Association of prediabetes-associated single nucleotide polymorphisms with microalbuminuria

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Abstract

Increased glycemic exposure, even below the diagnostic criteria for diabetes mellitus, is crucial in the pathogenesis of diabetic microvascular complications represented by microalbuminuria. Nonetheless, there is limited evidence regarding which single nucleotide polymorphisms (SNPs) are associated with prediabetes and whether genetic predisposition to prediabetes is related to microalbuminuria, especially in the general population. Our objective was to answer these questions. We conducted a genomewide association study (GWAS) separately on two population-based cohorts, Ansung and Ansan, in the Korean Genome and Epidemiology Study (KoGES). The initial GWAS was carried out on the Ansung cohort, followed by a replication study on the Ansan cohort. A total of 5682 native Korean participants without a significant medical illness were classified into either control group (n = 3153) or prediabetic group (n = 2529). In the GWAS, we identified two susceptibility loci associated with prediabetes, one at 17p15.3-p15.1 in the GCK gene and another at 7p15.1 in YKT6. When variations in GCK and YKT6 were used as a model of prediabetes, this genetically determined prediabetes increased microalbuminuria. Multiple logistic regression analyses revealed that fasting glucose concentration in plasma and SNP rs2908289 in GCK were associated with microalbuminuria, and adjustment for age, gender, smoking history, systolic blood pressure, waist circumference, and serum triglyceride levels did not attenuate this association. Our results suggest that prediabetes and the associated SNPs may predispose to microalbuminuria before the diagnosis of diabetes mellitus. Further studies are needed to explore the details of the physiological and molecular mechanisms underlying this genetic association.

Introduction

Diabetes mellitus (DM) and its vascular complications have become global socioeconomic and public health problems [1, 2]. Diabetic kidney disease (DKD), one of the most common microvascular complications of DM, seems to increase the risk of cardiovascular mortality [3, 4].
Thus, early identification of potential risk factor(s) of DKD and a preventive strategy against DKD are crucial for improvement of long-term health and survival.

Prediabetes, which refers to a plasma glucose level that is above the normal range but not high enough to meet the diagnostic criteria of DM, usually indicates a risk of conversion to type 2 DM (T2D) [5–7]. Even though not all patients with prediabetes progress to full-blown T2D, recent epidemiological studies have shown that subjects with prediabetes have various forms of vascular complications associated with T2D before the diagnosis of DM, which are also associated with an increased risk of kidney disease and cardiovascular morbidity and mortality [6–10]. Such findings suggest that even prediabetes may be a leading cause of complications that are typically attributed to DM.

Microalbuminuria, small amounts of albumin leakage into urine, indicates dysfunction of the glomerular filtration barrier, which is not only the early feature of a diabetic microvascular complication but also an independent risk factor of cardiovascular disease, even in nondiabetic populations [11–13]. In addition to the reports about the association between prediabetes and microalbuminuria, there have been many studies that reveal genetic variations associated with susceptibility to proteinuria in patients with T2D [14–19]. Such findings suggest that a complex interaction of genetic and environmental factors may have positive or negative influence(s) on both hyperglycemia and the related complications. Nonetheless, there is only limited evidence showing how genetic and nongenetic determinants of prediabetes may interact with microalbuminuria. Our aim was to clarify the association of prediabetes with microalbuminuria in the general population. Therefore, we conducted a genomewide association study (GWAS), which yielded useful results.

Results

The relation between prediabetes and microalbuminuria

The characteristics of each cohort and the study design are shown in Table 1 and S1 Fig. Out of the 5682 people included in the study, 2529 subjects had a diagnosis of prediabetes on the basis of fasting plasma glucose, 2-hour glucose in the oral glucose tolerance test, and glycated hemoglobin (HbA1c). The anthropometric, clinical, and laboratory details of the study participants—who were classified into two groups according to whether they had prediabetes—are shown in Table 2. Only urinary albumin-to-creatinine ratios (UACR) were log-transformed because all the biomarkers except UACR appeared to be normally distributed. Mean fasting glucose, mean postprandial glucose, and HbA1c in the prediabetic group were all higher than those in subjects with a normal glucose level. In comparison with the controls, an increase UACR was prominent in subjects with prediabetes (2.4 ± 0.6 vs. 2.5 ± 0.7, P < 0.0001 after log transformation). Participants in the prediabetic group were older and more obese, had higher blood pressure (BP), and a worse lipid profile than the control group did. In particular, a mildly increased C-reactive protein (CRP) concentration in plasma and a decreased kidney function were observed in the prediabetic group.

Table 1. Characteristics of the study population.

| Stage       | Study     | Sample type | Source                        | Number of samples | Males (%) | Age (years) |
|-------------|-----------|-------------|-------------------------------|-------------------|-----------|-------------|
| GWAS        | Ansung    | Control     | Korean Biobank Network        | 1362              | 589 (43)  | 52.6 ± 8.8  |
|             |           | Prediabetic | Korean Biobank Network        | 1092              | 494 (45)  | 55.5 ± 8.6  |
| Replication | Ansan     | Control     | Korean Biobank Network        | 1791              | 874 (49)  | 46.4 ± 6.2  |
|             |           | Prediabetic | Korean Biobank Network        | 1437              | 761 (53)  | 48.7 ± 7.6  |
| Combined    | Ansung + Ansan | Control     | Korean Biobank Network        | 3153              | 1463 (46) | 49.1 ± 8.0  |
|             |           | Prediabetic | Korean Biobank Network        | 2529              | 1690 (50) | 51.7 ± 8.7  |

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In the multiple linear regression analysis with adjustment for age, gender, body mass index and waist circumference, log-UACR had a strong positive relation with fasting blood glucose levels ($\beta = 0.0040$, $r^2 = 0.1257$, $P = 0.0479$) and HbA1c ($\beta = 0.1274$, $r^2 = 0.1276$, $P = 0.0059$). Furthermore, with the trend of increasing prevalence of microalbuminuria with age ($P$ for the trend $<0.0001$), we found that the prediabetic state was related to increased prevalence of microalbuminuria especially in younger subjects (Fig 1 and data not shown).

**GWAS on prediabetes**

We analyzed genetic data from 1092 participants in a prediabetic state (case) and 1362 healthy people (control) from the Ansung cohort. We carried out logistic regression analysis for prediabetes by including age and gender as covariates, and calculated the minimal p value for three genetic models (additive, recessive, and dominant). After a standard quality control procedure, we obtained genotyping results on 1,198,063 SNPs and generated a quantile-quantile plot (S2 Fig). The genomic inflation factor $\lambda$ was 1.01654 in the quantile-quantile plot, indicating minimal evidence of population stratification. The association analysis revealed that a total of 20

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**Table 2. Baseline characteristics grouped according to case-control status.**

| Characteristic                  | Control (n = 3153) | Case (n = 2529) | p    |
|--------------------------------|-------------------|----------------|------|
| Age (year)                     | 49.1 ± 8.0        | 51.7 ± 8.7     | <0.0001 |
| Gender (male, %)               | 1463 (46)         | 1255 (50)      | 0.0156 |
| Systolic BP (mmHg)             | 116.4 ± 12.1      | 118.5 ± 12.1   | <0.0001 |
| Diastolic BP (mmHg)            | 79.0 ± 9.7        | 80.4 ± 9.6     | <0.0001 |
| Body mass index (kg/m$^2$)     | 23.9 ± 2.7        | 24.6 ± 3.1     | <0.0001 |
| Waist circumference (cm)       | 80.0 ± 8.3        | 82.3 ± 8.6     | <0.0001 |
| eGFR*, mL/(min·[1.73 m$^2$])   | 77.7 ± 12.6       | 76.7 ± 12.8    | 0.0044 |
| eGFR < 60 (n, %)               | 93 (3)            | 157 (6)        | <0.0001 |
| Hemoglobin (g/L)               | 135 ± 16          | 136 ± 16       | 0.1676 |
| Albumin (g/L)                  | 42.3 ± 3.1        | 42.6 ± 3.4     | 0.0042 |
| Fasting glucose (mmol/L)       | 4.45 ± 0.37       | 4.72 ± 0.52    | <0.0001 |
| Postprandial glucose (mmol/L)  | 5.63 ± 1.13       | 7.09 ± 1.81    | <0.0001 |
| Hemoglobin A1c (%)             | 5.33 ± 0.22       | 5.89 ± 0.19    | <0.0001 |
| Triglycerides (mmol/L)         | 1.57 ± 0.98       | 1.84 ± 1.21    | <0.0001 |
| HDL-cholesterol (mmol/L)       | 1.18 ± 0.26       | 1.16 ± 0.35    | 0.0300 |
| LDL-cholesterol (mmol/L)       | 2.85 ± 0.8        | 3.03 ± 0.85    | <0.0001 |
| C-reactive protein (nmol/L)    | 1.71 ± 3.52       | 2.29 ± 4.86    | <0.0001 |
| UACR (mg/[g Cr])               | 14.5 ± 16.6       | 16.6 ± 20.5    | 0.0104 |
| Log-UACR (log mg/[g Cr])       | 2.4 ± 0.6         | 2.5 ± 0.7      | <0.0001 |
| Smoking history (%)            |                   |                | <0.0001 |
|   • Never smoker               | 1932 (61)         | 1406 (56)      |      |
|   • Ex-smoker                  | 433 (14)          | 408 (16)       |      |
|   • Intermittent smoker        | 100 (3)           | 67 (3)         |      |
|   • Chain smoker               | 688 (22)          | 648 (26)       |      |

Results are expressed as mean ± SD or as frequencies (and proportions).

BP, blood pressure; eGFR, estimated glomerular filtration rate; LDL, low-density lipoprotein; Log-UACR, log-transformed urine albumin/creatinine ratio; Cr, creatinine.

*a* estimated using the Chronic Kidney Disease Epidemiology Collaboration (CKD-EPI) equation.

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SNPs from 17 distinct genomic regions were significantly associated with prediabetes in the Ansung cohort (p_{GWAS} ranging from $9.54 \times 10^{-5}$ to $4.29 \times 10^{-6}$, S1 Table).

**Replication stage and combined analysis.** We performed subsequent association analysis of genetic data from the Ansan cohort on the basis of the 20 aforementioned SNPs and found significant associations for SNPs rs2908289, rs1799884, and rs917793 (p_{Replication} ranging from $9.79 \times 10^{-5}$ to $8.97 \times 10^{-6}$, S2 Table). Finally, these top three SNPs were used for combined analysis, which showed that prediabetes was significantly associated with a genetic polymorphism at 7p15 in the three genetic models (rs2908289, p_{Meta} = $2.5 \times 10^{-7}$; rs1799884, p_{Meta} = $1.7 \times 10^{-7}$; rs917793, p_{Meta} = $4.6 \times 10^{-8}$; Table 3).

Because the SNPs of interest were located in a genomic region encoding, for example, POLM, AEBP1, MYL7, GCK, YKT6, and CAMK2B, we conducted an imputation analysis to characterize these loci (Fig 2). The regional association plots using genotyped and imputed data revealed that rs2908289 and rs1799884 are confined to regions around the GCK gene encoding glucokinase, and rs917793 is located in the YKT6 gene encoding YKT6 (v-SNARE homolog).

![Fig 1. Age-specific prevalence of microalbuminuria by prediabetic state.](#)

*calculated by the Cochran-Armitage test for a trend. **defined as fasting glucose between 5.6 and 7.0 mmol/L, postprandial glucose between 7.8 and 11.0 mmol/L, or glycated hemoglobin between 5.7% and 6.4%. ***defined as a urine albumin/creatinine ratio (mg/[g creatinine]) between 30 and 300.

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Table 3. Analysis of the association of the top three single nucleotide polymorphisms (SNPs) with prediabetes.

| dbSNP ID | Nearest gene | Genotype | Study | Genotype frequency (%) | Additive\* | Dominant\* | Recessive\* |
|----------|--------------|----------|-------|------------------------|-----------|------------|------------|
|          |              |          | 1/2/3 | state                  | OR p      | OR p       | OR p       |
| rs2908289 | GCK/LOC105375257 | AA/AG/ GG | Control | 69.3 28.0 2.7 | 1.40 4.5 \times 10^{-6} | 1.34 2.5 \times 10^{-7} | 1.82 4.3 \times 10^{-5} |
|          |              |          |       | Case                   | 62.8 32.4 4.8 |           |            |            |
| rs1799884 | GCK          | TT/TC/CC | Control | 69.1 28.1 2.8 | 1.41 4.2 \times 10^{-6} | 1.35 1.7 \times 10^{-7} | 1.82 4.3 \times 10^{-5} |
|          |              |          |       | Case                   | 62.6 32.6 4.8 |           |            |            |
| rs917793 | YKT6         | TT/TA/AA | Control | 65.6 30.4 4.0 | 1.35 2.3 \times 10^{-6} | 1.36 4.6 \times 10^{-8} | 1.67 4.8 \times 10^{-5} |
|          |              |          |       | Case                   | 58.5 35.1 6.3 |           |            |            |

OR, Odds ratio.

\*calculated by logistic regression analysis with age and gender as covariates.

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Prediabetes-associated SNPs and microalbuminuria

In Mendelian randomization, which is a form of instrumental variable regression analysis, the instrumental variable-estimated size of the effect of prediabetes on log-UACR was highly significant and consistent for all outcomes ($\beta = 0.015$ [95% confidence interval (CI) 0.014–0.016] for fasting plasma glucose, 0.065 [0.011–0.012] for postprandial glucose, and 0.216 [0.198–0.234] for HbA1c; Table 4). We performed subsequence analysis for possible associations between candidate genetic polymorphisms and the clinical and laboratory characteristics of all the participants. Multiple logistic analysis adjusted for age and gender indicated that fasting glucose levels and SNPs in the \textit{GCK} and \textit{YKT6} genes were associated with microalbuminuria, and further adjustment for conventional risk factors, such as age, gender, smoking history, systolic BP, waist circumference, and serum triglyceride level, showed that the genetic polymorphism at rs2908289 was significantly associated with microalbuminuria, especially in recessive models (Table 5).

To determine the effect of different genotypes of rs2908289 on the possible risk factors of microvascular complications, each genotype was assessed by analysis of covariance and distinguished by the least significant difference method (Table 6). In addition to laboratory
indicators of prediabetes, SNP rs2908289 showed a significant dose-dependent relation with microalbuminuria and serum triglyceride levels.

Discussion

In this study examining the association between prediabetes genetic predisposition and microalbuminuria, we demonstrated that SNP rs2908289 located in the GCK gene can predict the risk of renal microvascular complications associated with DM in the general population. Our results suggest that genetic components of the pathogenesis of prediabetes may be linked with susceptibility to diabetic vascular complications.

Chronic glycemic exposure is one of the important causes of micro- or macroangiopathy in patients with DM [20]. An increase in oxidative stress, followed by activation of various signaling pathways, has been reported to cause endothelial dysfunction and inhibition of vascular protective mechanisms in the hyperglycemic milieu [21, 22]. Some authors have argued that hyperglycemia-induced microvascular changes may be evident before the onset of DM, can be aggravated during the natural course of the disease, and eventually may be related to the increased risk of poor renal or cardiovascular outcomes [8–10]. Nonetheless, there is limited evidence supporting a direct relation between prediabetes and albuminuria as markers of diabetic microvascular complications [14–16]. In this study, we showed that there is a close association between the plasma glucose level and urinary excretion of albumin in the general population, suggesting that even prediabetes may play a possible role in the pathogenesis of microvascular complications. It was also found that the prediabetic group was associated with increased prevalence of microalbuminuria in young subjects in contrast to the control group, with the overall tendency of increasing microalbuminuria with age in both groups, in line with studies on microalbuminuria in the general population [23, 24]. Additionally, there were fewer people between the ages of 50 and 59 years in the prediabetes group than in the control group. This result is probably due to the exclusion of individuals with T2D in the study design.

In the initial GWAS analysis, we found that SNPs in the GCK gene are independently associated with prediabetes in the general population. GCK, the cytoplasmic subtype of hexokinase that catalyzes the initial step of several glucose metabolic pathways, acts as a gatekeeper for glucose-induced activation or deactivation of biological processes [22, 25]. Thus, regulation of its enzymatic activity is important for glucose homeostasis. In agreement with these data, some genetic studies have shown that mutations in the gene encoding GCK may be significantly associated with both hyperinsulinemic hypoglycemia and maturity onset diabetes of the young [26]. Moreover, subjects with mutations in the gene encoding GCK have lifelong mild
hyperglycemia from birth, and this variant is significantly associated with progression to prediabetes or diabetes among the subjects with normal glucose tolerance [27–30]. Given that most of the existing studies were not confined to generally healthy subjects, it has been difficult to conclude that this genetic variation is directly associated with isolated prediabetes. Our results indicate that genetic variation of candidate genes may be associated with the initiation of aberrant glucose homeostasis in the general population.

The polymorphisms in the genetic loci within 7p15.3-p15.1 (encoding GCK) were found to be independently associated with an increased risk of microalbuminuria in subsequent analyses. Notably, it appears that there are additive genetic effects of the genotype of rs2908289 in

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**Table 5. Multivariate logistic regression analysis of microalbuminuria**

| Variable                      | Model I OR 95% CI | Model II OR 95% CI | Model III OR 95% CI |
|-------------------------------|------------------|--------------------|---------------------|
| Systolic BP (mmHg)           | 1.012 1.012–1.044| 1.025 1.008–1.041  |                     |
| Diastolic BP (mmHg)          | 1.023 1.003–1.043| 1.019 0.999–1.039  |                     |
| Body mass index (kg/m²)      | 1.072 1.013–1.135| 1.054 0.992–1.120  |                     |
| Waist circumference (cm)     | 1.031 1.011–1.051| 1.026 1.005–1.047  |                     |
| Hemoglobin (g/L)             | 0.985 0.851–1.141|                   |                     |
| Albumin (g/L)                | 1.238 0.526–2.914|                   |                     |
| eGFR, mL/(min [1.73 m²])     | 0.995 0.873–1.135|                   |                     |
| Fasting glucose (mmol/L)     | 1.028 1.008–1.049| 1.026 1.005–1.048  | 1.030 1.007–1.054  |
| Postprandial glucose (mmol/L)| 1.003 0.998–1.009|                   |                     |
| Hemoglobin A1c (%)           | 1.472 0.894–2.424|                   |                     |
| Triglyceride (mmol/L)        | 1.002 1.001–1.003|                   |                     |
| HDL-cholesterol (mmol/L)     | 1.007 0.992–1.023|                   |                     |
| LDL-cholesterol (mmol/L)     | 1.010 0.998–1.015|                   |                     |
| C-reactive protein (nmol/L)  | 1.28 0.989–1.655 |                   |                     |
| Smoking (vs. non-smoker)     |                   |                    |                     |
| Ex-smoker                    | 1.334 0.653–2.727|                   |                     |
| Intermittent smoker          | 2.888 1.314–6.349|                   |                     |
| Chain smoker                 | 1.292 0.697–2.396|                   |                     |
| rs2908289 Additive model     | 1.340 1.021–1.759| 1.409 1.067–1.862  | 1.262 0.922–1.729  |
| rs2908289 Dominant model     | 1.214 0.874–1.686|                   |                     |
| rs2908289 Recessive model    | 2.875 1.531–5.399| 3.227 1.693–6.152  | 2.568 1.210–5.453  |
| rs1799884 Additive model     | 1.336 1.018–1.754| 1.404 1.063–1.855  | 1.258 0.919–1.723  |
| rs1799884 Dominant model     | 1.209 0.870–1.680|                   |                     |
| rs1799884 Recessive model    | 2.871 1.529–5.393| 3.223 1.691–6.144  | 2.567 1.209–5.451  |
| rs917793 Additive model      | 1.222 0.942–1.584|                   |                     |
| rs917793 Dominant model      | 1.148 0.831–1.587|                   |                     |
| rs917793 Recessive model     | 1.937 1.072–3.498| 2.107 1.156–3.840  | 1.849 0.920–3.714  |

Model I: adjusted for age and gender.
Model II: adjusted for age, gender, smoking history, and serum triglyceride levels.
Model III: adjusted for age, gender, smoking history, systolic BP, waist circumference, and serum triglyceride levels.
OR, Odds ratio; CI, confidence interval.
*defined as a UACR between 30 and 300.

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the GCK gene on log-UACR. In a study on 42 families with maturity onset diabetes of the young 2 conducted in France, it was found that proteinuria develops at a relatively low frequency of 6% among these individuals, and another study showed similar results [27, 31]. On the other hand, functional and structural types of renal damage develop in GCK knockout mice according to several experimental studies [32, 33]. In a study on the Japanese population, GCK regulatory protein polymorphism was found to be significantly associated with the risk of chronic kidney disease (CKD) [34]. Recently, Böger et al. found that a common variant of the gene encoding a GCK-regulatory protein, a major cellular determinant of GCK enzymatic activity, may have a protective effect against T2D and CKD [35, 36]. Furthermore, Li et al. demonstrated that downregulation of GCK can contribute to the development of diabetic cardiomyopathy via increased oxidative stress and insulin resistance in an experimental animal model. Lastly, Szoja et al. reported that flow-mediated vasodilatation of the brachial artery is decreased in GCK mutation carriers [37, 38]. Such findings suggest that genetic variation in the GCK gene may exert positive or negative effect(s) not only on glucose metabolism in the liver and pancreas but also on the cardiovascular protective mechanism. Along with these studies, our findings of additive effects of the genotype of rs2908289 in the GCK gene on microalbuminuria should be considered in the future research on genetic interactions in complex trait variation.

Table 6. Association of different genotypes of rs2908289 with microalbuminuria and its risk factors.

| Variable                        | AA (n = 3762) | AG (n = 1705) | GG (n = 204) | p^a |
|---------------------------------|--------------|--------------|--------------|-----|
|                                 | Mean^a       | 95% CI^a     | Mean^a       | 95% CI^a     | Mean^a       | 95% CI^a     |
| eGFR, mL/(min [1.73 m^2])       | 77.1         | 76.7–77.4    | 77.7         | 77.2–78.2    | 77.0         | 75.5–78.5    | 0.1197       |
| Fasting glucose (mmol/L)        | 4.54         | 4.52–4.56    | 4.58         | 4.56–4.61    | 4.62         | 4.55–4.70    | 0.0052       |
| Postprandial glucose (mmol/L)   | 6.23         | 6.17–6.29    | 6.34         | 6.25–6.43    | 6.54         | 6.27–6.82    | 0.0179       |
| Hemoglobin A1c (%)              | 5.50         | 5.49–5.51    | 5.53         | 5.51–5.55    | 5.62         | 5.56–5.67    | <0.0001      |
| Log-UACR (log mg/g Cr)          | 2.42         | 2.37–2.46    | 2.45         | 2.34–2.51    | 2.61         | 2.44–2.79    | 0.0272       |

^aestimated using analysis of covariance after adjustment for age, gender, smoking history, systolic BP, waist circumference, and serum triglyceride levels, and their differences were estimated by the least significant difference method.

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In this study, we found that a novel SNP of isolated prediabetes, rs917793 in the YKT6 gene, is related to the development of microalbuminuria. YKT6 is a soluble N-ethylmaleimide-sensitive factor attachment protein receptor (SNARE) that participates in the production and release of exosomes, such as synaptic-vesicle exocytosis. YKT6 expression may be associated with the survival of cancer cells in lung and breast tissue [39, 40], but how genetic polymorphisms of the YKT6 gene may exert positive or negative effects on glucose homeostasis and on related complications is still poorly understood.

There are some limitations of this study. First, to efficiently analyze the potential nongenetic factors associated with hyperglycemia or microalbuminuria and to obtain consistent results, we excluded many participants with DM or other chronic medical diseases related to microalbuminuria. Such a study design resulted in a relatively small sample size for the GWAS. Nevertheless, considering that this study was community-based and that no participants had any medical illness, the sample size in this GWAS is large enough to show a relation between microalbuminuria and genetic factors of prediabetes. Second, a social desirability bias cannot be ruled out because the medical history and the use of medication, tobacco, or alcohol were all self-reported by the subjects. This approach may have contributed to the discrepancies with other studies. Third, although we performed urinary dipstick tests, we could not completely
rule out an asymptomatic urinary tract infection because urine culture analysis was not performed. Fourth, the test for albuminuria was conducted only during the first year of the study; therefore, it was impossible to obtain information about how many microalbuminuric patients progress to macroalbuminuria and DKD. Finally, because of population differences in allelic heterogeneity, generalization of the findings to all populations remains uncertain. Nonetheless, genetic studies on prediabetes in Asian populations may not necessarily confirm the same arrangement of susceptibility genes as those in other ethnic populations.

Despite the limitations, our results suggest that prediabetes may have a genetic impact on the development of diabetic complications. Therefore, genetic testing and early diagnosis of prediabetes can enable accurate prediction of diabetic complications, long-term health monitoring such as UACR for the subjects at a genetic risk and their families, and a preventive strategy against inevitable consequences such as DKD and end-stage renal disease. Furthermore, the genetic information associated with prediabetes can be useful for designing personalized treatments and for the development of new drugs for more precise medical care.

In conclusion, the results of the present study show a significant association between prediabetes and genetic polymorphisms of the GCK gene. Our results also revealed that SNP rs2908209 in the GCK gene can predict a lifelong risk of renal microvascular complications associated with prediabetes. We believe that genetic epidemiological studies of such associations may help to uncover the genetic basis of hyperglycemia-associated complications. Further research on the genetic factors influencing the development of albuminuria would be worthwhile because early detection and management of at-risk patients should help to inhibit the development and progression of CKD.

Materials and methods

Study population

This cross-sectional study on two population-based cohorts from Ansan and Ansung, Korea, was conducted by the Korean National Institute of Health as part of the KoGES, a Korean government-funded epidemiological survey to investigate trends in chronic diseases [41, 42]. All the participants volunteered and provided written informed consent prior to their enrollment. All the participants’ records, excluding the survey date and home region, were anonymized and deidentified before analysis by the authors. This study’s protocol was approved by the Institutional Review Board (IRB) of the Korea Centers for Disease Control and Prevention (IRB: 2014-10CON-06-P-E).

A total of 10,038 individual participants were examined biannually using laboratory tests, electrocardiograms, chest X-rays, and health questionnaires, and a 10-year follow-up study was recently completed. An oral glucose tolerance test was performed using the blood samples collected after fasting and 120 min after ingestion of 75 g of glucose. To clearly identify possible risk factor(s) for microvascular complications, the participants with a history of DM, essential hypertension, cancer, previously known kidney disease, or a UACR > 300 mg/(g creatinine) were excluded from this study (S1 Fig). Participants with urinary tract infection were excluded from the study on the basis of urinalysis. According to the 2016 American Diabetes Association standards of medical care, prediabetes is defined as aberrant fasting glucose (fasting plasma glucose between 5.55 and 6.94 mmol/L), impaired glucose tolerance (2-hr plasma glucose between 7.77 and 11.04 mmol/L), or HbA1c in the range 5.7–6.4% [5]. All eligible participants were subdivided into two groups: healthy controls and participants with prediabetes. Microalbuminuria was measured using spot morning urine and defined as a UACR between 30 and 300 mg/(g creatinine).
Genotyping
We analyzed the data on SNPs using publicly available whole-genome data from the Korea association resource (KARE) project from KoGES, and used the Affymetrix Genome-Wide Human SNP Array 5.0 (Affymetrix Inc., Santa Clara, CA, USA) to genotype the samples from the Ansan and Ansung cohorts. The Bayesian robust linear model with the Mahalanobis distance algorithm was used to determine the genotypes of each SNP. SNPs were excluded if any of the following criteria were met: (1) a call rate lower than 95%, (2) a minor allele frequency below 0.05, or (3) a significant deviation from Hardy–Weinberg equilibrium below 0.001. Among the SNPs filtered by these criteria, only tagging SNPs were used for analysis here.

Statistical analysis
Results are expressed as mean ± SD or as frequencies (and proportions). The normality of the distribution of parameters was analyzed by the Kolmogorov-Smirnov test. If a variable did not follow a normal distribution, a natural logarithm transformation was applied before statistical analysis. Because distribution of UACR was skewed to the right, this variable was log-transformed for further statistical analysis. Student’s t test was used to evaluate the differences in means between the two groups, or one-way analysis of variance was used for more than two groups. Categorical variables were assessed using chi-square analysis with Fisher’s exact test when the number of data points was small.

We performed linkage disequilibrium analysis of the analyzed polymorphisms of susceptibility genes using the Haploview software version 4.1 and generated a regional association analysis using HapMap, a web-based tool for identification and annotation of proxy SNPs. Mendelian randomization was applied to examine the direction of causality between microalbuminuria and the candidate SNP of susceptibility to prediabetes. We compared the results from the instrumental variable estimates of the association between genetic variation(s) and phenotypic measures to those from the least-squares linear regression using the Durbin–Wu–Hausman statistic [43, 44]. Odds ratios (ORs) with 95% CIs were calculated using multiple logistic regression models according to UACR (control vs. microalbuminuria). To identify age-adjusted effects of significant and suggestive SNPs, we performed analysis of covariance using the least significant difference method. All statistical analyses were conducted in the PLINK software, version 1.09, or Statistical Analysis Software (version 9.3; SAS Institute Inc., Cary, NC, USA).

Supporting information
S1 Fig. A flow chart of the study group enrollment process.
(TIF)

S2 Fig. A quantile plot of the observed p values (black dots) for the association of all 1,198,063 SNPs at the GWAS stage.
(TIF)

S1 Table. Genotype distribution of the top 20 single nucleotide polymorphisms (SNPs) associated with a prediabetic state in the Ansung cohort.
(DOC)

S2 Table. Genotype distribution of SNPs associated with a prediabetic state in the Ansan cohort.
(DOC)
Author contributions

Conceptualization: CHL JSP.
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Visualization: JSP JWC.
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References

1. Wild S, Roglic G, Green A, Sicree R, King H. Global prevalence of diabetes: estimates for the year 2000 and projections for 2030. Diabetes Care. 2004; 27(5): 1047–1053. Epub 2004/04/28. PMID: 15111519
2. Sarwar N, Gao P, Seshasai SR, Robin R, Kaptoge S, Di Angelantonio E, et al. Diabetes mellitus, fasting blood glucose concentration, and risk of vascular disease: a collaborative meta-analysis of 102 prospective studies. Lancet. 2010; 375(9733): 2215–2222. Epub 2010/07/09. doi: 10.1016/S0140-6736(10)60484-9 PMID: 20609967
3. Krolewski AS, Bonventre JV. High risk of ESRD in type 1 diabetes: new strategies are needed to retard progressive renal function decline. Semin Nephrol. 2012; 32(5): 407–414. Epub 2012/10/16. doi: 10.1016/j.semnephrol.2012.07.002 PMID: 23062980
4. Packham DK, Alves TP, Dwyer JP, Atkins R, de Zeeuw D, Cooper M, et al. Relative incidence of ESRD versus cardiovascular mortality in proteinuric type 2 diabetes and nephropathy: results from the DIA-METRIC (Diabetes Mellitus Treatment for Renal Insufficiency Consortium) database. Am J Kidney Dis. 2012; 59(1): 75–83. Epub 2011/11/05. doi: 10.1053/j.ajkd.2011.09.017 PMID: 22051245
5. American Diabetes Association. 2. Classification and Diagnosis of Diabetes. Diabetes care. 2016; 39 (Suppl 1): S13–S22.
6. Forouhi NG, Luan J, Hennings S, Wareham NJ. Incidence of Type 2 diabetes in England and its association with baseline impaired fasting glucose: the Ely study 1990–2000. Diabetic Med. 2007; 24(2): 200–207. Epub 2007/01/30. doi: 10.1111/j.1464-5491.2007.02068.x PMID: 17257284
7. Nathan DM, Davidson MB, DeFronzo RA, Heine RJ, Henry RR, Pratley R, et al. Impaired fasting glucose and impaired glucose tolerance: implications for care. Diabetes Care. 2007; 30(3): 753–759. Epub 2007/03/01. doi: 10.2337/dc07-9920 PMID: 17327355
8. Plantinga LC, Crews DC, Coresh J, Miller ER 3rd, Saran R, Yee J, et al. Prevalence of chronic kidney disease in US adults with undiagnosed diabetes or prediabetes. Clin J Am Soc Nephrol. 2010; 5(4): 673–682. Epub 2010/03/27. doi: 10.2215/CJN.07891109 PMID: 20338960
9. Grundy SM. Pre-diabetes, metabolic syndrome, and cardiovascular risk. J Am Coll Cardiol. 2012; 59 (7): 635–643. Epub 2012/02/11. doi: 10.1016/j.jacc.2011.08.080 PMID: 22322078
10. Ford ES, Zhao G, Li C. Pre-diabetes and the risk for cardiovascular disease: a systematic review of the evidence. J Am Coll Cardiol. 2010; 55(13): 1310–1317. Epub 2010/03/27. doi: 10.1016/j.jacc.2009.10.060 PMID: 20338491
11. Endemann DH, Schiffrin EL. Endothelial dysfunction. J Am Soc Nephrol. 2004; 15(8): 1983–1992. Epub 2004/07/31. doi: 10.1097/01.ASN.0000132474.50966.DA PMID: 15284284
12. Won JC, Lee YJ, Kim JM, Han SY, Noh JH, Ko KS, et al. Prevalence of and factors associated with albuminuria in the Korean adult population: the 2011 Korea National Health and Nutrition Examination
1. Dupuis J, Langenberg C, Prokopenko I, Saxena R, Soranzo N, Jackson AU, et al. New genetic loci
2. Matschinsky FM. Regulation of pancreatic beta-cell glucokinase: from basics to therapeutics. Diabetes.
3. Ohn JH, Kwak SH, Cho YM, Lim S, Jang HC, Park KS, et al. 10-year trajectory of beta-cell function and
4. Stratton IM, Adler AI, Neil HA, Matthews DR, Manley SE, Cull CA, et al. Association of glycaemia with
5. Steele AM, Shields BM, Wensley KJ, Colclough K, Ellard S, Hattersley AT. Prevalence of vascular com-
6. Calmettes G, Ribalet B, John S, Korge P, Ping P, Weiss JN. Hexokinases and cardioprotection. J Mol Cell Cardiol. 2015; 78: 107–115. Epub 2014/09/30. doi: 10.1016/j.yjmcc.2014.09.020 PMID: 25264175
7. de Jong PE, Hillege HL, Pinto-Sietsma SJ, de Zeeuw D. Screening for microalbuminuria in the general population: a tool to detect subjects at risk for progressive renal failure in an early phase? Nephrol Dial Transplant. 2003; 18(1): 10–13. PMID: 12480951
8. Konta T, Hao Z, Abiko H, Ishikawa M, Takahashi T, Ikeda A, et al. Prevalence and risk factor analysis of microalbuminuria in Japanese general population: the Takahata study. Kidney Int. 2006; 70(4): 751–756. Epub 2006/06/28. doi: 10.1038/sj.ki.5001504 PMID: 16807548
9. Kim WH, Lee JW, Suh YH, Hong SH, Choi JS, Lim JH, et al. Exposure to chronic high glucose induces beta-cell apoptosis through decreased interaction of glucokinase with mitochondria: downregulation of glucokinase in pancreatic beta-cells. Diabetes. 2005; 54(9): 2602–2611. Epub 2005/08/27. doi: 10.2337/diabetes.54.9.2602
10. Osbак KK, Colclough K, Saint-Martin C, Beer NL, Bellanne-Chantelot C, Ellard S, et al. Update on mutations in glucokinase (GCK), which cause maturity-onset diabetes of the young, permanent neonatal diabetes, and hyperinsulinemic hypoglycaemia. Hum Mutat. 2009; 30(11): 1512–1526. Epub 2009/10/01. doi: 10.1002/humu.21116 PMID: 19790256
11. Steele AM, Shields BM, Wensley KJ, Colclough K, Ellard S, Hattersley AT. Prevalence of vascular complications among patients with glucokinase mutations and prolonged, mild hyperglycaemia. JAMA. 2014; 311(3): 279–286. Epub 2014/01/17. doi: 10.1001/jama.2013.283980 PMID: 24430320
12. Ohn JH, Kwak SH, Cho YM, Lim S, Jang HC, Park KS, et al. 10-year trajectory of beta-cell function and insulin sensitivity in the development of type 2 diabetes: a community-based prospective cohort study. Lancet Diabetes Endocrinol. 2016; 4(1): 27–34. Epub 2015/11/19. doi: 10.1016/S2213-8587(15)00336-8 PMID: 26577716
13. Matschinsky FM. Distribution of pancreatic beta-cell glucokinase: from basics to therapeutics. Diabetes. 2002; 51(Suppl 3): S394–S404. Epub 2002/12/12.
14. Dupuis J, Langenberg C, Prokopenko I, Saxena R, Soranzo N, Jackson AU, et al. New genetic loci implicated in fasting glucose homeostasis and their impact on type 2 diabetes risk. Nat Genet. 2010; 42(2): 105–116. Epub 2010/01/19. doi: 10.1038/ng.520 PMID: 20081858
31. Velho G, Blanche H, Vaxillaire M, Bellanne-Chantelot C, Pardini VC, Timsit J, et al. Identification of 14 new glucokinase mutations and description of the clinical profile of 42 MODY-2 families. Diabetologia. 1997; 40(2): 217–224. doi: 10.1007/s001250050666 PMID: 9049484

32. Gu Y, Mao Y, Li H, Zhao S, Yang Y, Gao H, et al. Long-term renal changes in the liver-specific glucokinase knockout mouse: implications for renal disease in maturity-onset diabetes of the young 2. Transl Res. 2011; 157(3): 111–116. Epub 2011/02/15. doi: 10.1016/j.trsl.2011.04.003 PMID: 21316027

33. Xu W, Li H, Wang R, Lei Z, Mao Y, Wang X, et al. Differential expression of genes associated with the progression of renal disease in the kidneys of liver-specific glucokinase gene knockout mice. Int J Mol Sci. 2013; 14(3): 6467–6486. Epub 2013/03/23. doi: 10.3390/ijms14036467 PMID: 23519111

34. Hishida A, Takashima N, Turin TC, Kawai S, Wakai K, Hamajima N, et al. GCK, GCKR polymorphisms and risk of chronic kidney disease in Japanese individuals: data from the J-MICC Study. J Nephrol. 2014; 27(2): 143–149. Epub 2014/02/19. doi: 10.1016/s0022-5371(14)30003-0 PMID: 24535998

35. Bocher CA, Gorski M, Li M, Hoffmann MM, Huang C, Yang Q, et al. Association of eGFR-Related Loci Identified by GWAS with Incident CKD and ESRD. PLoS Genet. 2011; 7(9): e1002292. Epub 2011/10/08. doi: 10.1371/journal.pgen.1002292 PMID: 21980298

36. Brouwers MC, Jacobs C, Bast A, Stehouwer CD, Schaper NC. Modulation of glucokinase regulatory protein: A double-edged sword? Trends Mol Med. 2015; 21(10): 583–594. Epub 2015/08/04 PMID: 26432016

37. Li H, Wang X, Mao Y, Hu R, Xu W, Lei Z, et al. Long term liver specific glucokinase gene defect induced diabetic cardiomyopathy by up regulating NADPH oxidase and down regulating insulin receptor and p-AMPK. Cardiovasc Diabetol. 2014; 13: 24. Epub 2014/01/23. doi: 10.1186/1475-2840-13-24 PMID: 24447392

38. Szopa M, Osmenda G, Wilk G, Matejko B, Skupien J, Zapala B, et al. Intima-media thickness and endothelial dysfunction in GCK and HNF1A-MODY patients. Eur J Endocrinol. 2015; 172(3): 277–283. Epub 2014/12/17. doi: 10.1530/EJE-14-0713 PMID: 25501962

39. Ruiz-Martinez M, Navarro R, Marrades RM, Vinolas N, Santasusagna S, Munoz C, et al. YKT6 expression, exosomal release, and survival in non-small cell lung cancer. Oncotarget. 2016. Epub 2016/06/11.

40. Ooe A, Kato K, Noguchi S. Possible involvement of CCT5, RGS3, and YKT6 genes up-regulated in p53-mutated tumors in resistance to docetaxel in human breast cancers. Breast Cancer Res Treatment. 2007; 101(3): 305–315. Epub 2006/07/06.

41. Kim Y, Han BG. Cohort Profile: The Korean Genome and Epidemiology Study (KoGES) Consortium. Int J Epidemiol. 2016. Epub 2016/04/17.

42. Hong KW, Jin HS, Lim JE, Cho YS, Go MJ, Jung J, et al. Non-synonymous single-nucleotide polymorphisms associated with blood pressure and hypertension. J Hum Hypertens. 2010; 24(11): 763–774. Epub 2010/02/12. doi: 10.1038/jhh.2010.9 PMID: 20147996

43. Smith GD, Ebrahim S. ‘Mendelian randomization’: can genetic epidemiology contribute to understanding environmental determinants of disease? Int J Epidemiol. 2003; 32(1): 1–22. Epub 2003/04/12. PMID: 12869998

44. Kivimaki M, Smith GD, Timpson NJ, Lawlor DA, Batty GD, Kahnem M, et al. Lifetime body mass index and later atherosclerosis risk in young adults: examining causal links using Mendelian randomization in the Cardiovascular Risk in Young Finns study. Eur Heart J. 2008; 29(20): 2552–2560. Epub 2008/06/14. doi: 10.1093/eurheartj/ehn252 PMID: 18550552