Strong Gravitational Lens Statistics using the \textit{Herschel}-ATLAS

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2 May 2014

ABSTRACT

The identification of strong gravitational lenses in large surveys has historically been a rather time consuming exercise. Early data from the \textit{Herschel} Astrophysical Terahertz Large Area Survey (\textit{Herschel}-ATLAS) demonstrate that lenses can be identified efficiently at submillimetre wavelengths using a simple flux criteria. Motivated by that development, this work considers the statistical properties of strong gravitational lens systems which have been, and will be, found by the \textit{Herschel}-ATLAS. Analytical models of lens statistics are tested with the current best estimates for the various model ingredients. These include the cosmological parameters, the mass function and the lens density profile, for which we consider the singular isothermal sphere (SIS) and the Navarro, Frenk & White (NFW) approximations. The five lenses identified in the \textit{Herschel}-ATLAS Science Demonstration Phase suggest a SIS density profile is preferred, but cannot yet constrain $\Omega_\Lambda$ to an accuracy comparable with other methods. The complete \textit{Herschel}-ATLAS data set should be sufficient to yield competitive constraints on $\Omega_\Lambda$. Whilst this huge number of lenses has great potential for constraining cosmological parameters, they will be most powerful in constraining uncertainty in astrophysical processes. Further investigation is needed to fully exploit this unprecedented data set.

1 INTRODUCTION

Since the first strong gravitational lens was discovered in 1979 (Walsh, Carswell & Weymann), hundreds have been found and studied (Treu 2010). As well as serendipitous discoveries, dedicated surveys have searched for lens candidates. The Cosmic Lens All-Sky Survey (Browne et al. 2003) identified 22 new strong gravitational lenses from imaging in the radio waveband. The Sloan Lens ACS Survey (Bolton et al. 2006; Auger et al. 2009) has so far confirmed 85 new gravitational lenses using spectroscopic data from the SDSS to identify candidates for follow-up with high-resolution HST imaging. The Cosmological Evolution Survey (Faure et al. 2008; Jackson 2008) identified 67 lenses from visual inspection of HST imaging. In addition there are projects, such as MUSCLES (Jackson et al. 2011) and SQLS (Inada et al. 2012), that are looking for new lenses by reprocessing data from large surveys such as UKIDSS and SDSS.

There has been much interest in studying gravitational lenses because they magnify sources that would otherwise be too faint or distant to see. Historically these lens systems have been difficult to identify in sufficiently large numbers to be statistically useful. Some lens searches involve painstaking ‘by eye’ analysis of each object observed. Now new methodologies are being developed to improve the efficiency of finding lenses and to improve the conversion rate from candidate to confirmed lensed source. These include using colours (Ofek et al. 2007), spectroscopic redshifts (Auger et al. 2009), and other imaging processing techniques (Marshall 2008).

For a rigorous statistical study of gravitational lenses, they should all come from a single survey. This is because the probability of identifying lensed sources varies between surveys due to the different instrumental sensitivities and selection effects.

For sources observed in the submillimetre waveband it has long been thought that galaxy number counts should fall off rapidly with increasing flux density. This is because of the rapid evolution of proto-spheroidal galaxies into starburst galaxies (Granato et al. 2004). This drop in counts with flux using submillimetre observations was anticipated to yield an effective way of identifying lenses (Blain 1996; Negrello et al. 2007). It has since been successfully demonstrated using the \textit{Herschel}-ATLAS data (Negrello et al. 2011; González-Nuevo et al. 2012). Lensed sources are identifiable much more easily because the lensing magnification pushes them into higher flux densities which have intrinsically low source populations.

The \textit{Herschel} Space Observatory, launched in 2009, is the only space-based observatory to cover a spectral range from the far infrared to the submillimetre. It provides a unique window through which to study large scale structure. The \textit{Herschel} Astrophysical Terahertz Large Area Survey (Eales et al. 2010, \textit{Herschel}-ATLAS) will cover 550 deg$^2$ of the sky making it the \textit{Herschel} survey with the largest...
area. [Negrello et al. (2010) showed that strong gravitational lenses could be identified with almost 100% efficiency by selecting sources which have 500 μm flux densities greater than 100 mJy. González-Nuevo et al. (2012) further improved on this technique by refining the flux selection to incorporate other Herschel wavebands and optical/infrared data from the VIKING (Fieuren et al. 2012; Sutherland et al. 2012) survey to look for counterparts. This technique is forecast to yield ten times more lenses than that by Negrello et al. The expected efficiency of 70% is lower that that achieved by Negrello et al., but still impressive. Using these techniques, the Herschel-ATLAS team expect to find as many as a thousand strongly lensed candidates.

We emphasise that this article considers the likelihood of strong gravitational lenses which are defined as lensed systems with multiply imaged sources. Recent articles on the effect of lensing on the source number counts [Lima et al. 2011; Wardlow et al. 2012] do not consider strong lensing in the sense that sources are multiply imaged, just that they are highly magnified. Note, this results in the lens probability distribution assuming a NFW density profile becoming much more similar to that of the SIS density profile than is seen in our analysis. Looking to the future, ‘strongly’ lensed sources are being selected by their excess submillimetre flux, or effectively by their magnification, with limited optical follow-up to confirm multiple imaging. Since requiring optical follow-up may limit the number of confirmed sources, it may in future turn out to be statistically favourable to relax the multiple image requirement to gain a statistical advantage. However, our approach, looking at strong gravitational lens systems with multiple images, is used to be consistent with the nature of the sources identified by Negrello et al. (2010).

The statistical properties of gravitationally lensed sources were first considered in the 1980s [Turner, Ostriker & Gott 1984] and later applied as a possible way of constraining the cosmological model [Li & Ostriker 2002; Mitchell et al. 2005; Zhang, Cheng & Wu 2009; Dobke et al. 2009]. With the prospect of a thousand lenses being identified by Herschel-ATLAS, we take a fresh look at the analytical theory behind predicting strong gravitational lens statistics in Sections 2 and 3. In Section 4, the lenses identified by Negrello et al. (2010) are compared against the analytical predictions, and then in Section 5, we consider the parameter constraints possible with the full Herschel-ATLAS data set.

Unless otherwise stated we use cosmological parameters: \( \Omega_m = 0.27, \Omega_a = 0.73, \Omega_\Lambda = 0.046, h = 0.7, n = 0.97, \) and \( \sigma_8 = 0.81 \) [Komatsu et al. 2011].

## 2 Lens Statistics: The Theory

A source is defined to be strongly gravitationally lensed when it has been multiply imaged due to the gravitational effects of mass between the source and the observer. The probability of a source being lensed is known as the optical depth (\( \tau \)). It depends on three main components: the cosmology which determines the co-moving volume element at a given redshift; the normalised mass function (\( n \)) which describes the number density of halos; and the lensing cross-section (\( \sigma \)), which is the area in the lens plane where the separation between the lens and source is sufficiently small for strong lensing to occur:

\[
\frac{d\tau(z_s, z_l)}{dz_l} = \int_0^\infty \frac{d\mu}{dz_l} (1 + z_l)^3 n(m, z_l) \sigma(m, z_l) \, dm, \tag{1}
\]

where \( d\tau \) is the angular diameter distance to the lens, \( z_l \) is the lens redshift, \( z_s \) is the source redshift, and \( m \) is the virial mass of the lens [Schneider, Ehlers & Falco 1992]. This is explained in detail in the Appendix, where we describe how to calculate each of the different components. As well as the dependence on the lensing redshift \( z_l \), the dependence on other parameters such as the total magnification of the lensed source or the angular separation of the images can be calculated.

When comparing the theoretical prediction for the distribution of strong lenses with observational surveys, another factor that needs to be taken into account is magnification bias [Turner, Ostriker & Gott 1984].

### 2.1 Magnification Bias

Magnification bias is a correction to account for the fact that lensed sources are magnified and are consequently observed to be brighter than the population from which they are drawn. Therefore sources which are lensed come from a population with a lower flux. Since there are generally more fainter sources than there are brighter sources, magnification bias almost always increases the chance of finding a lens compared to if the lenses came from the same flux population at which they are observed.

The magnification bias at a particular flux is calculated as the ratio of the number of lenses which will be magnified to the flux over the number of objects which intrinsically have that flux times the probability of lensing:

\[
B(F) = \frac{\int_1^\infty \tau(\mu) N\left(\frac{\mu}{F}\right) d\mu}{\tau N(F)} = \frac{\int_1^\infty P(\mu) N\left(\frac{\mu}{F}\right) d\mu}{\tau N(F)}, \tag{2}
\]

where \( P(\mu) = \tau(\mu)/\tau \) is the probability of a magnification \( \mu \) given that the source has been lensed.\(^1\)

The overall bias is then calculated as the average of Eq. (2) over all fluxes:

\[
B = \lim_{F \to \infty} \int_1^\infty \int_1^\infty P(\mu) N\left(\frac{\mu}{F}\right) d\mu dF = \frac{\int_1^\infty N(F) dF}{\int_1^\infty F_{\text{lim}} dF}, \tag{3}
\]

where \( F_{\text{lim}} \) is the flux limit of the observational survey.

Generally the flux distribution is assumed to take the form of a simple power law \( N(F) = N_0 (F/F_0)^{-\beta} \) and so Eq. (3) then reduces to \( B = \int \tau(\mu) \mu^\beta d\mu \). However, for this particular survey the slope \( \beta \) is so steep that \( \mu^\beta \) grows faster than \( \tau(\mu) \) decreases and so this integral does not converge. But assuming that this power law approximation

\(^1\) Note that we do not here include a \( \mu^{-1} \) factor as in Turner, Ostriker & Gott (1984). As detailed in Fukugita & Turner (2007) since both \( \tau \) and \( \tau(\mu) \) are considered in the same plane this ‘correction’ factor is not required.
for $N(F)$ holds for infinitely small fluxes is clearly incorrect; there will come a point where the flux counts fall off. To account for this we introduce an exponential term that stops the growth of $N(F)$ below a given flux $F_c$ meaning that $N(F)$ will return to zero as $F$ tends to zero. We use $N(F) = N_c(F/F_c)^{-\beta} \exp(-(F/F_c)^2)$ and the magnification bias becomes:

$$B = \frac{\int_1^\infty P(\mu) \mu^2 \int_{F_{\text{lim}}}^\infty F^{-\beta} \exp(-(F/\mu F_c)^2) dF d\mu}{\int_{F_{\text{lim}}}^\infty F^{-\beta} \exp(-(F/F_c)^2) dF}.$$  \tag{4}

In this work we simplify the above expression by setting $F_{\text{lim}}$ to 0. Eq. \(4\) then reduces to:

$$B = \int_1^\infty P(\mu) \mu d\mu,$$  \tag{5}

which intuitively makes sense as it is just the average magnification. This work assumes that $F_{\text{lim}} = 0$.

Whilst Eq. \(1\) is relatively simple, it is more complex to implement in practice due to the uncertainties associated with each of the constituent components. In the rest of this section we explore the current best estimates for each of the components and consider the likely impact of uncertainty in each of them on the probability distributions for observing gravitational lenses.

### 3 LENS STATISTICS: DEMONSTRATING THE THEORY

In this section the different components of the lensing optical depth are considered. We discuss the current best estimates for each component and demonstrate the corresponding effects of uncertainty on the optical depth.

#### 3.1 Dark Energy Density

The community is, in general, confident in the $\Lambda$CDM cosmological model refined from CMB (Bennett et al. 2003), BAO (Percival et al. 2010) and supernovae (Riess et al. 1998) observations. However it is still worth considering how the different cosmological parameters would affect these lensing statistics. If variations in the underlying cosmology significantly affect the lensing optical depth then the statistics of gravitational lenses could provide another complementary test of the cosmology. On the other hand, if the cosmological parameters have very little impact on the lensing statistics then we can be confident that any assumptions we draw are not dependent on the accuracy of the $\Lambda$CDM assumption.

The dark energy density ($\Omega_\Lambda$) contributes to the lensing optical depth (described in Eq. \(1\) and the Appendix) by increasing the comoving volume the light from a source travels through. This makes it more likely that a source will encounter an object and be lensed. If we assume a flat universe, this increase in $\Omega_\Lambda$ is balanced by a reduction in the matter density ($\Omega_m$). This decrease in $\Omega_m$ will reduce the density perturbation growth rate and hence decrease the mean mass in the mass function. Figure \(1\) shows an example of the differential lensing optical depth as a function of lens redshift for $\Omega_\Lambda = 0.66$ (purple dashed line) and 0.7 (blue solid line). This assumes a source redshift $z_s = 2$, SIS lens density profile and a Tinker et al. (2008) mass function. The increase in the comoving volume element and decrease in the density perturbation growth rate combine so that the overall optical depth decreases with increasing $\Omega_\Lambda$.

Previous work has used results from the CLASS (Browne et al. 2003) survey to constrain $\Omega_\Lambda$. The CLASS data set, which includes 13 lenses, was used to constrain the total dark energy density $\Omega_\Lambda$ for a flat Universe (Chae et al. 2002; Mitchell et al. 2003) and the dark energy equation of state parameter (Zhang, Cheng & Wu 2004). For illustrative purposes, here we consider just varying the dark energy density for a flat universe. The shape of the differential lensing optical depth does not significantly change as a function of $\theta$ and $\mu$ (not shown).

![Figure 1. The differential lensing optical depth as a function of lens redshift for $\Omega_\Lambda = 0.66$ (purple dashed line) and 0.7 (blue solid line). This assumes a source redshift $z_s = 2$, SIS lens density profile and a Tinker et al. (2008) mass function. The increase in the comoving volume element and decrease in the density perturbation growth rate combine so that the overall optical depth decreases with increasing $\Omega_\Lambda$.](image)

#### 3.2 Lens Distributions

Understanding the lens distribution is important as lensing is a gravitational effect and can tell us about the distribution of all matter not just baryonic matter. The distribution of matter over-densities, or halos, is frequently approximated using an analytical approximation called the mass function.

The mass function $n(m, z)$ (sometimes written as $dn/dm$) predicts the differential number of halos of a given mass as a function of redshift. It is frequently approximated...
using the approach by Press & Schechter (1974):

\[
\frac{m^2 n(m, z_1)}{\bar{\rho}} \frac{dm}{m} = \nu f(\nu) \frac{d\nu}{\nu},
\]

where \(\bar{\rho}\) is the average matter density, \(\nu(m, z_1) = \delta^2/\sigma^2\), \(\delta_c\) is the critical density required for spherical collapse and \(\sigma^2\) is the variance in the linear density fluctuation field (see also Appendix A3). There have been a number of estimates proposed for \(f(\nu)\). Press & Schechter (1974) derived an approximation based on theoretical arguments. This was later shown to underpredict the abundance of halos of high/low mass relative to N-body simulations (Sheth & Tormen 1999). Improved estimates have been proposed (Jenkins et al. 2001; Warren et al. 2006), most recently by Tinker et al. (2008) whose fit introduced a specific dependence on redshift.

This was later shown to underestimate the abundance of high/low mass relative to N-body simulations (Sheth & Tormen 1999). Improved estimates have been proposed (Jenkins et al. 2001; Warren et al. 2006), most recently by Tinker et al. (2008) whose fit introduced a specific dependence on redshift.

So how does uncertainty in the mass function \(n(m, z_1)\) affect the differential lensing optical depth? Figure 2 shows the differential lensing optical depth for both the Press & Schechter (1974) and Tinker et al. (2008) approximations for the mass function. The Press & Schechter mass function estimates lower values of \(d\tau/dz_1\) than that the Tinker et al. mass function. The shape of \(d\tau/dz_1\) tends to favour: larger \(z_1\) for the Press & Schechter mass function compared to the Tinker et al. mass function; and smaller \(\theta\) for the Press & Schechter mass function compared to the Tinker et al. mass function. This is because the Press & Schechter mass function predicts a larger \(n\) at small \(m\), and a smaller \(n\) at high \(m\), compared to the Tinker et al. mass function. Integrating over all \(m\) results in a smaller \(d\tau/dz_1\) and \(d\tau/d\theta\) for the Press & Schechter mass function because the higher mass halos contribute most to the lensing optical depth. \(d\tau(\mu)/d\mu\) maintains the same shape for both mass functions (not shown).

Whilst it is theoretically possible to generate a lensed source with a lens made purely of dark matter, the way in which candidate lens sources are confirmed relies on follow-up imaging to identify a lens candidate. Without an observed lens candidate it is likely the source would be rejected as a potential lens. This is not something that has been accounted for in this work. The lensing optical depth is approximated with a mass function for the total matter distribution. This is worth noting as a potential source of error.

### 3.3 Halo Density Profile

Eq. 6 is most sensitive to the lens density profile component. The simplest model for a halo density profile that can explain the flat rotation curves of galaxies is a singular isothermal sphere (SIS). This has a density profile given by:

\[
\rho(r) = \frac{\sigma^2}{2\pi G r^2},
\]

where \(\sigma\) is the velocity dispersion (see Appendix A3). This profile has been successful in reproducing a number of features of gravitational lensing (Auger et al. 2009). However the profiles of dark matter halos generated via numerical simulations are better fit by the profile proposed by Navarro, Frenk & White (1997, NFW):

\[
\rho(r) = \frac{\rho_s}{\left(\frac{r}{r_s}\right)^2 \left(1 + \frac{r}{r_s}\right)},
\]

where \(r_s\) is the scale radius and \(\rho_s\) is the density at \(r_s\) (see Appendix A6). Neither the SIS or NFW density profiles is probably realistic for a lens made of a dark matter halo and baryon component. Here we use them for illustrative purposes to demonstrate the scale of uncertainty and its significant effect in the resulting optical depth.

SIS profiles produce more strong gravitational lenses than the NFW profile. This is because they have a steeper inner density gradient (\(\rho \propto r^{-2}\) as opposed to \(\rho \propto r^{-1}\) for the NFW) and steeper inner profiles are more effective at producing multiply imaged lenses (Li & Ostriker 2002). NFW profiles produce lenses with higher mean magnifications than those for the SIS. This is because lenses that are less concentrated (such as the NFW) require the source and lens to be very closely aligned for strong lensing to result. It is this close alignment that results in a higher magnification (Wyithe, Turner & Spergel 2001). In other words lenses that are more concentrated (such as the SIS) do not need the

![Figure 2. The differential lensing optical depth as a function of redshift for \(z_1 = 2\) and a SIS lens profile. The green dashed line is for the Press & Schechter mass function, which is compared to the Tinker et al. mass function (blue solid line). The Press & Schechter mass function underestimates the lensing optical depth compared to the Tinker et al. mass function.](image-url)
Strong Gravitational Lens Statistics using the Herschel-ATLAS

Figure 3 compares the differential optical depth for the SIS and NFW halo profiles. As well as demonstrating the known results for $d\tau/d\mu$ discussed above, other effects are seen for $d\tau/dz$ and $d\tau/d\theta$. The NFW profile tends to predict more small-separation lenses and fewer large-separation lenses than the SIS profile. Also, the mean lens redshift is much smaller for the NFW density profile than the SIS.

These results confirm that the theoretical strong gravitational lensing probabilities are highly sensitive to the density gradient near the centre of dark matter halos (Li & Ostriker 2002). This is an area of tension in current observations (Comerford et al. 2006; Shu et al. 2008). High resolution observations of galaxy rotation curves suggest halo density profiles with a flat central core. However, studies of gravitational lensing statistics suggest a steep, or ‘cuspy’, density profile for halo cores (Chen & McGaugh 2008).

A potential resolution for this is that the profiles of the gravitational lens population are biased with respect to the general population (Mandelbaum 2009); alternative profiles have also been suggested. Following Navarro, Frenk & White (1997), higher resolution numerical simulations (Moore et al. 1999) suggested that the central density gradient should lie somewhere in between that of the NFW ($\rho \propto r^{-1}$) and SIS ($\rho \propto r^{-2}$) profiles. It may be that parameterising the slope of the density profile with an Einasto (Retana-Montenegro & Frutos-Alfaro 2011) or a Séracse (Mandelbaum 2008; Elíasdóttir & Møller 2007) profile would be more appropriate.

Alternatively, CLASS found a number of lenses at low separation, but none at higher $\theta$, which suggested a combination of NFW and SIS density profiles. They proposed that if the mass of the lens is greater than a given cut-off $m_c$ then the lens density profile is approximated using a NFW approximation, else a SIS profile is used (Li & Ostriker 2002). This was not motivated by direct observations/simulations of halo profiles, but it has been shown to provide a reasonable fit to lensing statistics for a mass cut-off scale of $m_c = 10^{13} M_\odot h^{-1}$ (Zhang, Cheng & Wu 2009).

The NFW halo density profile could well be accurate but the halo concentration could be incorrect. The concentration of the halo (defined as $c = r_{\text{vir}}/r_h$) has been studied extensively over the last decade. On average, the median halo concentration decreases with increasing mass and redshift; this can be approximated using analytical expressions (Navarro, Frenk & White 1997; Bullock et al. 2001; Eke et al. 2001).

All these suggestions assume a spherically symmetric lens, but realistically the halos are more likely to be irregularly shaped with a spherically symmetric average. Some way of accounting for this variation will be required to provide a consistent explanation across dark matter simulations and lensing studies.

### 3.4 Source Distributions

The analytical calculation of the lensing optical depth in Eq. 1 only describes the probability of a source being lensed. It does not account for the probability of a source being at a particular redshift in the first place. So if you wanted to estimate the total optical depth this source redshift distribution needs to be measured observationally.

For the Herschel sources it is generally not possible to measure accurate spectroscopic redshifts due to the lack of emission lines in that band and the low angular resolution. Much effort has been put into deriving accurate source
redshift distributions for the Herschel source catalogues via other methods such as matching sources to optical counterparts with spectroscopic redshifts (Smith et al. 2011), CO line spectroscopy follow-up (Harris et al. 2012), and using models based on average galaxy SEDs (Amblard et al. 2010; Lapi et al. 2011; Pearson et al. 2012). All these methods have their own biases/constraints, so a potential test that could be done with this lens work would be to fix all the other components and investigate the source distribution.

In this section we have seen that the lensing statistics are sensitive to both astrophysical and cosmological parameters. The lensing optical depth is most sensitive to the lens density profile, both in magnitude and as a function of all $z_l$, $\mu$ and $\theta$. The mass function has little effect on the optical depth as a function of $\mu$, but does for $z_l$ and $\theta$. Of all the parameters considered, the optical depth was least sensitive to the dark energy density. It should be noted that before competitive constraints on the cosmological parameters can be achieved, the uncertainty in the properties of the lensing galaxies needs to be constrained.

4 LENS STATISTICS: HERSCHEL-ATLAS SDP

Five new lensed sources were discovered using the flux criteria technique on the Herschel-ATLAS Science Demonstration Phase (SDP) data (Negrello et al. 2010). Table 1 summarises the estimated source and lens redshifts for those lenses. Whilst this is probably not enough lenses with which to do convincing statistics, we can still consider whether or not these observations are consistent with the analytical models discussed in Section 3.

Here we consider both the absolute differential optical depth as a function of lensing redshift $d\tau(z_l)/dz_l$ and the conditional differential optical depth $\mu$.

\[
P(\mu) = \frac{1}{\tau} \frac{d\tau(\mu)}{d\mu} \quad \text{and} \quad P(z_l) = \frac{1}{\tau} \frac{d\tau(z_l)}{dz_l},
\]

which describe the probability of a lens being at a certain $\mu$ or $z_l$ given that lensing has occurred. We do not consider $P(\theta)$ here because estimates are not yet available for the image separations of the Herschel-ATLAS SDP lenses.

By choosing to consider the conditional probability, rather than the absolute probability, the calculation is simplified. The magnification bias (see Section 2.1) does not need to be calculated as the conditional probability is effectively the ‘normalised’ optical depth for fixed $z_l$. The conditional probability is therefore the expected distribution of the lenses as a function of $\mu$ or $z_l$ given that lensing has occurred.

We start by looking to constrain the lens density profile since this is the component to which the statistics of strong lenses appeared to be most sensitive in Section 3.3. We start by considering the conditional probability as a function of magnification $P(\mu)$ and redshift $P(z_l)$ to constrain the lens density profile between the NFW and SIS.

4.1 Constraining the Lens Density Profile with $P(\mu)$

Magnifications are available for the Herschel-ATLAS SDP lenses ID081 and ID130 (Negrello et al. 2010). Theoretical distributions of the conditional probability as a function of magnification were calculated as described in the Appendix. They were calculated assuming a Tinker et al. mass function, and a standard ΛCDM cosmology. The source and lens redshifts are as per the estimates in Table 1. The resulting $P(\mu)$ are shown in Figure 4 for the two different estimates of the lens density profile (SIS and NFW), along with the two observational estimates of the magnifications for ID081 and ID130. This shows that the magnification of ID081 source is within the 99% confidence interval of the theoretical probability found for both the SIS and NFW density profiles. The magnification for the ID130 source is consistent with the model with the SIS density profile at a 95% level, but inconsistent with the model with the NFW profile at a 99% level. These results appear to favour the SIS density profile.

One way of quantifying the combination of the ID081 and ID130 results to choose between the NFW and SIS options is to use a likelihood. The likelihood of the data given the model parameters is defined as:

\[
\mathcal{L}(\mu_i | \rho_{SIS/NFW}) = \prod_{i=1}^{N_1} P(\mu_i),
\]

where $N_1 = 2$ is the number of lensed sources being considered. To give the most favourable result for each of the lens models, the highest probability associated with the range of observed magnification is chosen for each $P(\mu_i)$. Calculating this gives $\mathcal{L}(\mu_i | \rho_{SIS}) = 3 \times 10^{-3}$ and $\mathcal{L}(\mu_i | \rho_{NFW}) < 5 \times 10^{-6}$ which confirms that a SIS density profile is preferred overall.

The observational estimates for the magnification in Figure 4 are consistently higher than those preferred by the SIS lens density profile. We note that magnifications close to 1 will be much harder to identify observationally than higher magnifications. Therefore we would expect that the observational values would be biased against magnifications close to $\mu = 1$. This suggests that the SIS likelihood is actually higher. Overall we see the Herschel-ATLAS SDP lens magnifications show that the SIS density profile is much preferred over the NFW density profile.

| ID | $z_s$ | $z_l$ | $\mu$ |
|----|------|------|------|
| 9  | 0.673±0.057 | 1.577±0.008 | – |
| 11 | 0.793±0.0012 | 1.786±0.005 | – |
| 17 | 0.943±0.0009 | 2.308±0.011 | – |
| 81 | 0.299±0.0002 | 3.042±0.001 | 25±7 |
| 130| 0.220±0.002 | 2.626±0.0003 | 6±1 |

Table 1. Estimates of the source ($z_s$) and lens ($z_l$) redshifts and magnification ($\mu$) for the five lenses discovered in the Herschel-ATLAS SDP (Negrello et al. 2010). The quoted errors correspond to a 68% confidence interval.
The SDP lenses. Again, magnification (i.e., standard ΛCDM cosmology, for any angular separation and the Appendix assuming a Tinker et al. mass function and a impacts

In Section 3 we saw that the lens density profile strongly magnification given lensing has occurred, for source and lens redshifts in Table 1. Solid blue lines use the SIS lens density profile; dashed red lines use the NFW lens density profile. Shading indicates 68% confidence intervals. Solid green vertical lines indicate the estimated magnification of gravitational lenses ID081 and ID130 found by Negrello et al. (2010).

4.2 Constraining the Lens Density Profile with $P(z_i)$

In Section 3 we saw that the lens density profile strongly impacts $P(z_i)$ and $P(\theta)$ as well as $P(\mu)$. Data is not yet available for the angular separation of the lensed images, therefore we just look at $P(z_i)$ here. There were only two data points available for the lens magnifications but lens redshift estimates are available for all of the Herschel-ATLAS SDP lenses.

Figure 5 shows $P(z_i)$ for all five of the Herschel-ATLAS SDP lenses. Again $P(z_i)$ was calculated as described in the Appendix assuming a Tinker et al. mass function and a standard ΛCDM cosmology, for any angular separation and magnification (i.e. $\theta > 0$ and $\mu > 1$). The source redshifts are as per the estimates in Table 1.

Three of the five Herschel-ATLAS lensed sources (ID009, ID011, and ID017) show lens redshifts most consistent with 68% confidence limits for the SIS density profile model. These positions are still consistent with the NFW model at around the 95% confidence level. The other two Herschel-ATLAS lensed sources (ID081 and ID130) are both at much smaller lens redshifts. The $z_i$ are consistent with both the SIS and NFW models again at a 95% confidence level.

To quantify these results we can again calculate likelihood as:

$$L(z_i | \rho_{\text{SIS/NFW}}) = \prod_{i=1}^{N_i} P(z_i).$$

Calculating this gives $L(z_i | \rho_{\text{SIS}}) = 0.014$ and $L(z_i | \rho_{\text{NFW}}) = 0.026$. So, in contrast to what was shown with $P(\mu)$, the NFW profile is preferred over the SIS density profile.

Again we should note that there is likely to be an observational bias here. The way in which these lens systems are selected requires a lens object to be identified observationally. Objects which are closer will tend to be brighter and therefore easier to identify. Therefore it is likely that there is a selection bias for nearer lenses which has not been taken into account. This bias favours the NFW profile, so if it were taken into account the SIS profile would likely have a more favourable likelihood.

4.3 Constraining the Dark Energy Density with $d\tau/dz_i$

Here we consider whether the dark energy density can be competitively constrained using the lens redshifts $z_i$ of the Herschel-ATLAS SDP lenses as compared to the theoretical $d\tau(z_i)/dz_i$. The tests above provide conflicting evidence for the SIS and NFW lens density profiles. Later in Section 5 we will put these results into context using simulations, and project how many lensed sources are needed for a definitive result. Meanwhile, for illustrative purposes, here we assume a SIS density profile. Similarly since there is no information on the angular separation, we make the assumption that the mass function is well fit by a Tinker et al. mass function.

Now we consider absolute probabilities we also need to take into account the other sources that were not lensed. This is to account for the probability of lensing occurring as well as the distribution of the lenses as a function of $z_i$. The likelihood is given by:

$$L(z_i | \Omega_\Lambda) = \prod_{i=1}^{N_i} \frac{d\tau(z_i | z_i)}{dz_i} \prod_{j=1}^{N_u} (1 - \tau(z_{u,j})),$$

where $N_u$ is the number of unlensed sources. Given an estimated redshift distribution $\mathcal{N}(z_i)$ then, because the optical depth is small, we can approximate $L(z_i | \Omega_\Lambda)$ as (Mitchell et al. 2003):

$$\ln L(z_i | \Omega_\Lambda) = \sum_{i=1}^{N_i} \ln \frac{d\tau(z_i | z_i)}{dz_i} + \sum_{j=1}^{N_u} \ln(1 - \tau(z_{u,j}))$$

$$= \sum_{i=1}^{N_i} \ln \frac{d\tau(z_i | z_i)}{dz_i} - \sum_{j=1}^{N_u} \tau(z_{u,j})$$

$$= \sum_{i=1}^{N_i} \ln \frac{d\tau(z_i | z_i)}{dz_i} - \int \mathcal{N}(z_i) \tau(z_u) dz_u.$$
Figure 5. Conditional probability of finding a lens at a certain lens redshift given lensing has occurred, for all five Herschel-ATLAS lenses. Source and lens redshifts are in Table 1. Solid blue lines use the SIS lens density profile; dashed red lines use the NFW density profile. Shading indicates 68% confidence intervals. Solid green vertical lines indicate the estimated lens redshifts of the gravitational lenses found by Negrello et al. (2010).

Herschel-ATLAS SDP lenses but $z_s$ in the second term is a discrete range from the minimum to maximum redshifts of Herschel-ATLAS sources observed.

The redshift distribution of the Herschel-ATLAS sources is quite uncertain; for illustrative purposes here we assume that all the sources are at a redshift of 2. We take the number of unlensed sources as 252 from Clements et al. (2010); the resulting $\Omega\Lambda$ is invariant to small changes in $N_u$. Figure 6 shows the resulting likelihood as a function of the dark energy density (assuming a flat universe). This corresponds to an estimate of $\Omega\Lambda = 0.81^{+0.02}_{-0.04}$. This is consistent with the current best estimate for $\Omega\Lambda$ at the 95% level.
To summarise, there are some conflicting results, which is not surprising given the current limited number of confirmed gravitational lensed sources. In the next section we put these results in context using simulations. We also consider what constraints can be achieved with a larger number of lenses.

5 LENS STATISTICS: FULL HERSCHEL-ATLAS SURVEY

So far just five strong gravitational lenses have been confirmed in the Herschel-ATLAS SDP data. With the full data set it should be possible to extract enough lensed sources with which to do more meaningful statistical analyses (Negrello et al. 2010; González-Nuevo et al. 2012). Here we consider whether the full set of 100-1000 lenses will be sufficient to constrain the different models for the density profile, mass function and ΩΛ discussed above.

The methodology is described in more detail below, but here is a general overview. The differential optical depth was calculated as described in the Appendix using different density profiles, mass functions and ΩΛ. A thousand sets of Nl lenses were generated by randomly sampling the differential optical depth at different redshifts zi, magnifications µ, and angular separations θ. The corresponding likelihoods were then calculated as described in Equations 9, 10 and 11 for each of the sets. The set size Nl is varied to see how successfully the different approximations for the ΩΛ dependence realised from random sampling of theoretical probability distributions P(µ) for both SIS and NFW profiles. The two distributions only slightly overlap so it is fairly easy to distinguish between SIS and NFW using this approach - it will be even easier with more results.

The actual likelihoods found from the SDP data above in Section 4 were log L(zi|P(SIS)) = −2.5 and log L(µ|P(NFW)) < −5.3. These are both consistent with the bottom plot which gives likelihoods for lens redshifts sampled from P(µ) assuming a SIS density profile.

The blue and red curves are for likelihoods calculated with SIS and NFW density profiles respectively. The two distributions only slightly overlap in the first instance, and not at all in the second case (the NFW distribution is at smaller Ω than on the plot). It is easy to distinguish between SIS and NFW density profiles because their likelihood are distinct even with this small amount of data.

The actual likelihoods found from the SDP data above in Section 4 were log L(zi|P(SIS)) = −2.5 and log L(µ|P(NFW)) < −5.3. These are both consistent with the bottom plot which gives likelihoods for lens redshifts sampled from P(µ) assuming a SIS density profile.

Whilst in Section 4 P(µ) was found to be more consistent with a SIS model than a NFW model, P(zi) suggested the opposite. Here the same test as above was performed again but this time sampling the lens redshift as opposed to the magnification. Figure 5 shows the results. The top plot is for sets of Nl = 2 lensed sources (as we have for the current SDP sample). The top half shows the distribution of likelihoods found using SIS and NFW probability for a thousand sets for both the SIS and NFW lens profiles as described in Eq. 3. So in all four different combinations were considered. For each combination, the number counts for a range of different likelihoods were calculated. These results are shown in Figure 7.

The blue and red curves are for likelihoods calculated with SIS and NFW density profiles respectively. The two distributions only slightly overlap in the first instance, and not at all in the second case (the NFW distribution is at smaller Ω than on the plot). It is easy to distinguish between SIS and NFW density profiles because their likelihood are distinct even with this small amount of data.

The actual likelihoods found from the SDP data above in Section 4 were log L(zi|P(SIS)) = −2.5 and log L(µ|P(NFW)) < −5.3. These are both consistent with the bottom plot which gives likelihoods for lens redshifts sampled from P(µ) assuming a SIS density profile.
Figure 8. Normalised likelihoods realised from random sampling of theoretical probability distributions $P(z_i)$ given a Tinker et al. mass function and standard ΛCDM cosmological parameters, for SIS (blue solid line) and NFW (red dashed line) density profiles. The top plot is for sets of $N_1 = 5$ lensed sources (as we have for the current SDP sample); the lower plot is for $N_1 = 20$ lenses. The top half shows the distribution of likelihoods found using SIS and NFW probabilities for lenses randomly sampled from the NFW $P(\mu)$. Similarly for the bottom plot but here the lenses are randomly sampled from the SIS $P(\mu)$. The shading represents 68% and 95% confidence intervals. The two distributions overlap for $N_1 = 5$ so it is less easy to distinguish between SIS and NFW profiles; but at $N_1 = 20$ they are fairly easy to distinguish.

less easy to distinguish between SIS and NFW using $P(z_i)$ than when using $P(\mu)$. Likelihoods found above in Section 4 gave $\log L(z_i | \rho_{\text{SIS}}) = -1.9$ and $\log L(z_i | \rho_{\text{NFW}}) = -1.6$. Although the NFW likelihood is higher we see from Figure 8 that is not inconsistent with a SIS lens. More data is needed to properly distinguish between them; $N_1 = 20$ lenses would be sufficient as seen from the bottom plot.

Therefore to summarise, whilst the results for the density profile appeared conflicted initially in Section 4 here we see that actually they are both most consistent with a SIS density profile.

Values of the angular separation of lensed images are not yet available, but here we consider whether using $\theta$ would allow us to constrain the best fit mass function. Following the same approach as above but now sampling $\theta$ from $P(\theta)$, the results for $N_1 = 100$ lenses are shown in Figure 9.

This shows that even with angular separations for 100 lensed sources it would not be possible to distinguish between the Press & Schechter and Tinker et al. mass functions. Of the order 500 sources would be required as a minimum. The number of detections is expected to be this high, but the shear amount of follow-up that would be required to get enough angular separations $\theta$ makes this approach unfeasible.

Finally, we consider constraining the dark energy density $\Omega_\Lambda$ with the absolute differential optical depth $d\tau/d\bar{z}_1$. The lens redshifts are randomly sampled from the optical depth assuming the standard ΛCDM cosmological parameters. Rather than considering two alternatives of $\Omega_\Lambda$, here we consider a range of $\Omega_\Lambda$ to calculate the likelihoods. Therefore rather than comparing the normalised distributions of $L(d\tau/d\bar{z}_1)$, instead we take the mean value $\bar{L}$ for each $\Omega_\Lambda$ and this is plot in Figure 6. The Tinker et al. mass function and SIS density profile are used. The plot is for sets of $N_1 = 5$ (as we have for the current SDP sample), 100 and 1000 lensed sources. The shaded areas are 68% limits.

We see that the SDP estimate of $\Omega_\Lambda$ is consistent with the sort of errors expected for a sample of 5 lens redshifts, but with ~100 lens redshifts the error on $\Omega_\Lambda$ should be of order 0.01 i.e. comparable to the current best methods.

6 DISCUSSION AND CONCLUSIONS

Until recently the number of strong gravitational lenses that had been identified was relatively small for any particular survey. Now this situation is about to change dramatically as future surveys are expected to yield numbers of strong...
More uncertain and further work is clearly required to understand how this and other other uncertainties can most effectively be constrained.

Here we have seen that the dominant uncertainties in gravitational lens statistics are not cosmological but astrophysical i.e. understanding the nature of the lensing objects. Whilst more investigation is required, we can see that new data sets of strong gravitational lenses will be fundamental for constraining both halo astrophysics as well as cosmological parameters.

**ACKNOWLEDGMENTS**

Many Thanks to Loretta Dunne and Steve Maddox for useful discussions. For the purposes of this work Jo Short is supported by STFC STEP fellowship, Elizabeth Pearson received funding from a STFC studentship, Peter Coles is partly supported by STFC Standard Grant ST/G002231/1, and Steve Eales and Peter Coles are supported by STFC Rolling Grant ST/H001530/1.

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Figure 10. Mean likelihood as a function of dark energy density realised from random sampling of theoretical probability distributions $d\sigma/dz$ given a SIS density profile and [Tinker et al.](2008) mass function. The shading represents 68% confidence intervals. Assuming the other astrophysical uncertainties can be constrained, we see that competitive constraints on $\Omega_\Lambda$ can be achieved with the full Herschel-ATLAS data set.
APPENDIX A: THEORETICAL STATISTICS OF GRAVITATIONAL LENSES

A1 General Formulation

A source is defined to be strongly gravitationally lensed when it has been multiply imaged due to the gravitational effects of mass between the source and the observer. The probability of a source being lensed is described by the optical depth ($\tau$). It depends on three main components: the cosmology which determines the co-moving volume element at a given redshift; the normalised mass function ($n$) which describes the number density of halos; and the lensing cross-section ($\sigma$), which is the area in the lens plane where the separation between the lens and source is sufficiently small for strong lensing to occur:

$$\tau(z_s) = \int_0^{z_s} \frac{dD_l}{dz_l} (1 + z_l)^3 n(m, z_l) \sigma(m, z_l) \, dm \, dz_l,$$  \hspace{1cm} (A1)

where $m$ is the mass of the lens, $z_l$ is the redshift of the lens, $z_s$ is the redshift of the source, and $D_l$ is the angular diameter distance from the observer to the lens (Schneider, Ehlers & Falco 1993). Below we describe how each of these components can be calculated.

A2 Angular Diameter Distance

The angular diameter distance between the two redshifts $z_a$ to $z_b$ is given by:

$$D_{ab} = \frac{c}{H_0} \frac{1}{1 + z_b} \int_{z_a}^{z_b} E(z) \, dz,$$  \hspace{1cm} (A2)

where $E(z) = \Omega_m (z=0)(1 + z)^3 + \Omega_k (z=0)]^{0.5}$.

A3 Mass Function

The differential number density of halos ($n$, sometimes written as $dn/dm$) is commonly estimated using this expression derived by Press & Schechter (1974):

$$\frac{m^2 n(m, z_l)}{\rho} \frac{dm}{m} = \nu f(\nu) \frac{d\nu}{\nu},$$  \hspace{1cm} (A3)

$\bar{\rho} = \Omega_m \rho_c$ is the average matter density, $\nu(m, z_l) = \delta_c^2/\sigma^2(m, z_l)$, $\delta_c = 1.686$ is the critical density required for spherical collapse and $\sigma^2$ is the variance in the linear density fluctuation field. Two approximations for $f(\nu)$ are those by Press & Schechter (1974) and Tinker et al. (2008):

$$f_{PS}(\nu) = \sqrt{\frac{1}{2\pi\nu}} \exp\left(-\frac{\nu}{2}\right),$$  \hspace{1cm} (A4)

and

$$f_T(\nu) = A \left(\frac{\delta_c}{b \nu}\right)^{-\alpha} + 1 \exp\left(-\frac{\nu c}{\delta_c^2}\right),$$  \hspace{1cm} (A5)

where $A = 0.186(1 + z)^{-0.14}$, $a = 1.47(1 + z)^{-0.06}$, $b = 2.57(1 + z)^{-\alpha}$, $\alpha = 1.068$, and $c = 1.19$.

\footnote{This definition assumes a flat universe.}

\footnote{We use the Bardeen et al. (1986) transfer function to calculate the linear power spectrum.}

A4 Lensing Cross-section

The gravitational lensing cross-section ($\sigma$) is the area in the lens plane where the lens and source are within the critical angular separation $\beta_c$ for strong lensing to occur (Schneider, Ehlers & Falco 1993):

$$\sigma = \pi \beta_c^2 D_l^2.$$  \hspace{1cm} (A6)

The critical angular separation $\beta_c$ that produces $\sigma$ depends on the shape of the lens density profile. Two different approximations for the lens density profile (SIS and NFW) are considered in Sections A5 and A6. First the general approach for spherically symmetric lens density profiles is reviewed.

The critical angle ($\beta_c$ in Eq. A6) can be estimated using the lens equation:

$$\beta = \theta - \alpha,$$  \hspace{1cm} (A7)

where $\alpha = \hat{\alpha} D_{ls}/D_s$ (see Figure A1). Note this assumes a flat space-time so that the Euclidean relation $\text{separation} = \text{angle} \times \text{distance}$ holds.

The lens equation can alternatively be written in terms of the radial distances, $\eta/D_s = \xi/D_l - \hat{\alpha} D_{ls}/D_s$. This can be simplified further using dimensionless distances $y = \eta/\eta_0$ and $x = \xi/\xi_0$ where $\eta_0 = \xi_0 D_s/D_l$, to give the lens equation of the form:

$$y = x - \alpha \frac{D_{ls}}{D_s \xi_0}.$$  \hspace{1cm} (A8)

Using the thin lens approximation, which uses the fact that the lens is thin compared to the total path traveled by the light, the deflection angle $\hat{\alpha}$ can be calculated as:

$$\hat{\alpha}(x) = \frac{4G m(x)}{x \rho_0 c^2} ,$$  \hspace{1cm} (A9)

where $m(x) = 4\pi \int_0^x \rho(r) \, dr \, dz$ and $r^2 + s^2 = z^2 + x^2$.

The distance in the lens plane subtended by the critical...
angle is simply \( y_c = \beta_s D_s / y_0 \). This is called the critical radius and is calculated as the value of \( y(x) \) where \( dy/dx = 0 \). For values of \( y \) below \( y_c \), multiple images are formed. We will see below that two multiple images are formed in the case of the SIS density profile, and three in the case of the NFW density profile.

### A5 Singular Isothermal Sphere (SIS) density profile

The SIS density profile is given by:

\[
\rho(r) = \frac{v^2}{2\pi G r^2}. \tag{A10}
\]

The velocity dispersion \( \sigma \) can be approximated as:

\[
v(m, z_i) = 92.3 \left[ \Delta v_{\text{vir}}(z_i) E(z_i) \frac{m}{10^{14} h^{-1} M_\odot} \right]^{1/3} \text{km/s}, \tag{A11}
\]

where \( \Delta v_{\text{vir}}(z_i) = 18\pi^2 + 60[\Omega(z_i) - 1] - 32[\Omega(z_i) - 1]^2 \) and \( \Omega(z_i) = \Omega_0 [(1+z_i)^3 / E(z_i)^2] \) (Mitchell et al. 2003).

The deflection angle (Eq. A9) becomes \( \alpha = 4\pi v^2 / c^2 \) and the lens equation (Eq. A8) reduces to:

\[
\xi_0 = \frac{4\pi v^2 D_s D_\ell}{c^2} \tag{A12}
\]

The critical value of \( y \) below which strong lensing (or multiple images) occurs is \( y_c = 1 \). We see from Figure A2 that two images are generated in this case. The lensing cross-section is given by:

\[
\sigma_{\text{SIS}} = \pi\beta_s^2 D_\ell^2 = \frac{\pi y_c^2 D_\ell^2}{D_s^2} = \pi \left[ \frac{4\pi v^2 D_s D_\ell}{c^2 D_\ell^2} \right]^2. \tag{A13}
\]

Other properties of the lenses, such as the magnification and angular separation, can also be calculated. The probability that a lens will be magnified more than \( \mu \) is included in the cross-section as:

\[
\sigma(\mu) = \sigma A(\mu), \tag{A14}
\]

where \( A(\mu) \) describes the fractional area satisfying magnifications greater than \( \mu \):

\[
A(\mu) = \frac{1}{\pi y_c^2} \int_{-y_c}^{y_c} \pi y \Theta[\mu(y) - \mu] \, dy. \tag{A15}
\]

where \( \Theta \) is a step function. The magnification of an image is given by:

\[
\mu(y) = \frac{y \, dy}{x \, dx}^{-1}. \tag{A16}
\]

Schneider, Ehlers & Falco (1993) and since the solutions of the SIS lens equation are \( x_{\pm} = y \pm 1 ((y < 1) \), this becomes:

\[
\mu_{\pm}(y) = \frac{1}{y} \pm 1. \tag{A17}
\]

So the total magnification \( \mu(y) = \mu_{-}(y) + \mu_{+}(y) = 2/y \). Substituting this into Eq. A15 gives:

\[
A(\mu) = \frac{1}{\pi y_c^2} \int_{-y_c}^{y_c} \pi y \Theta[2/y - \mu] \, dy = \frac{4}{\mu^2} \text{ for } \mu > 2. \tag{A18}
\]

Therefore \( \sigma_{\text{SIS}}(\mu) = 4\sigma_{\text{SIS}} \mu^{-2} \) for \( \mu > 2 \).

Similarly the probability that a lens will have images with an angular separation greater than \( \theta \) is included in the cross-section as:

\[
\sigma_{\text{SIS}}(\theta > \theta) = \sigma_{\text{SIS}} \Theta(\vartheta - \theta). \tag{A19}
\]

The separation between the two images can be seen from the lens equation to be \( \Delta x = 2 \). The angular separation becomes \( \vartheta = \xi_0 \Delta x / D_\ell = 8\pi v^2 D_s / c^2 D_s \).

### A6 Navarro, Frenk & White (NFW) density profile

Navarro, Frenk & White (1997) proposed a halo density profile of the form:

\[
\rho(r) = \frac{\rho_s}{\left( \frac{r}{r_s} \right) \left( 1 + \frac{r}{r_s} \right)^2}, \tag{A20}
\]

where \( r_s \) is the scale radius and \( \rho_s \) is the density at \( r_s \). The scale radius \( r_s \) can be calculated as \( r_s = r_{\text{vir}} / c \) where the virial radius \( r_{\text{vir}} \) is defined as the radius at which the mean density of the mass is 200 times that of the matter density. The concentration can be approximated using (Bullock et al. 2001):

\[
c(m, z_i) = \frac{9}{1+z_i} \left( \frac{m}{1.5 \times 10^{14} M_\odot} \right)^{-0.13}, \tag{A21}
\]

and \( \rho_s \) can be calculated using, \( m = 4\pi \rho_s r_s^3 \left[ \ln(1 + c) - c/(1 + c) \right] \). The lens equation (Eq. A8) becomes:

\[
y = x - \frac{\mu y}{x} g(x), \tag{A22}
\]
where $\xi_0 = r_s$, $\mu_s = 4\rho_s r_s / \Sigma_{\text{crit}}$. The critical surface density is given by:

$$\Sigma_{\text{crit}} = c^2 D_s / 4\pi G D_l D_{ls}, \quad (A23)$$

and:

$$g(x) = \int_0^x u \int_0^\infty \left( (u^2 + v^2)^{-0.5} (1 + (u^2 + v^2)^{0.5}) \right)^{-2} dv \, du. \quad (A24)$$

It can be seen from Figure A2 that the critical value of $y$ below which strong lensing (or multiple images) occurs is when $dy/dx = 0$. So the cross-section is given by:

$$\sigma_{\text{NFW}}(z_s, z_l, m) = \pi \beta_c D_l^2 = \pi \left( \frac{y_c \eta_0}{D_s} \right) D_l^2 = \pi y_c^2 r_s^2. \quad (A25)$$

As for the SIS profile, the dependence of $\sigma$ on other factors such as magnification and angular separation can be calculated. The lensing cross-section can be calculated in the same way as described in Equations A14 and A15. In the case of the NFW profile this does not reduce into a simple expression as is the case for the SIS. Instead $\sigma_{\text{NFW}}(> \mu)$ needs to be calculated using the full expression detailed in Equations A14 to A16. This calculation is extensive and a good approximation is available from Oguri et al. (2002).

Again to calculate the dependence of the cross-section of the angular separation, $\sigma_{\text{NFW}}(> \theta) = \sigma_{\text{NFW}} \Theta(\vartheta - \theta)$. The separation between the two furthest images can be seen from the lens equation to be $\Delta x = 2x_c$ where $x_c = \beta_c D_s$. The angular separation is therefore $\vartheta = \xi_0 \Delta x / D_l = 2r_s x_c / D_l$. 

