A Decentralized Approach to Bayesian Learning

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Abstract

Motivated by decentralized approaches to machine learning, we propose a collaborative Bayesian learning algorithm taking the form of decentralized Langevin dynamics in a non-convex setting. Our analysis show that the initial KL-divergence between the Markov Chain and the target posterior distribution is exponentially decreasing while the error contributions to the overall KL-divergence from the additive noise is decreasing in polynomial time. We further show that the polynomial-term experiences speed-up with number of agents and provide sufficient conditions on the time-varying step-sizes to guarantee convergence to the desired distribution. The performance of the proposed algorithm is evaluated on a wide variety of machine learning tasks. The empirical results show that the performance of individual agents with locally available data is on par with the centralized setting with considerable improvement in the convergence rate.

1 Introduction

With the recent advances in computational infrastructure, there has been an increase in the use of larger machine learning models with millions of parameters. Even though there is a parallel increase in the size of training datasets for these models, there is a significant disparity between the amount of existing data and the data required to train the large models to avoid overfitting and provide good generalization performance. Such models trained in point estimate settings such as Maximum A Posteriori (MAP) neglect any associated epistemic uncertainties and make overconfident predictions. Bayesian learning framework provides a principled way to avoid over-fitting and model uncertainties by estimating the posterior distribution of the model parameters. However, analytical solutions of exact posterior or sampling from the exact posterior is often impossible due to the intractability of the evidence. Therefore, one needs to resort to
approximate Bayesian methods such as Markov Chain Monte Carlo (MCMC) sampling techniques. To this effect, we focus on a specific class of MCMC methods, called Langevin dynamics to sample from the posterior distribution and perform Bayesian machine learning.

Langevin dynamics derives motivation from diffusion approximations and uses the information of a target density to efficiently explore the posterior distribution over parameters of interest \[1\]. Langevin dynamics, in essence, is the steepest descent flow of the relative entropy functional or the KL-divergence with respect to the Wasserstein metric \[2\]. Just as the gradient flow converges exponentially fast under a gradient-domination condition, Langevin dynamics converges exponentially fast to the stationary target distribution if the relative entropy functional satisfies the log-Sobolev inequality \[3\].

The Unadjusted Langevin Algorithm (ULA) is a popular inexact first-order discretized implementation of the Langevin dynamics without an acceptance/rejection criteria. Analysis of convergence properties of the ULA and other Langevin approximations has been a topic of active research over past several years \[6\]. Reference \[3\] shows that a bias exists in the ULA for any arbitrarily small (fixed) step size, even for a Gaussian target distribution. Controlling the bias and exponential convergence of KL divergence for strongly log-concave smooth target distributions using ULA is discussed in \[3,6,7,9,11\]. Non-asymptotic bounds on variation error of the Langevin approximations for smooth log-concave target distributions have been established by \[6\] and \[8\]. Assuming a Lipschitz continuous Hessian, \[6\] introduces a modified version of the Langevin algorithm requiring fewer iterations to achieve the same precision level. Tight relations between the Langevin Monte Carlo for sampling and the gradient descent for optimization for (strongly) log-concave target distributions are presented in \[7\].

Similarly, using the notion of gradient flows over probability space and KL-divergence, \[9\] analyzes the non-asymptotic convergence of discretized Langevin diffusion. These results were improved and extended with particular emphasis on scalability of the approach with dimension, smoothness, and curvature of the function of interest in \[10,11\].

Compared to log concave ULA settings where local properties replicate the global behavior and optimal values are attained in a single pass, non-convex objective functions naturally require multiple passes through training data. Analysis of ULA in such cases often requires assuming that the negative log of the target distribution satisfies some dissipative property \[13,17\], contractivity condition \[18\], or limiting the non-convexity to a local region \[19,20\]. In particular, \[13\] makes the first attempt in analyzing non-asymptotic convergence in a nonconvex setting and shows SGLD tracks continuous Langevin diffusion in quadratic Wasserstein distance for empirical risk minimization. Recent work \[14,19,21\] reports computational efficiency of sampling algorithm to optimization methods in the nonconvex setting. The approach is extended to relaxed dissipativity conditions, to evaluate dependent data streams and provides sharper convergence estimates uniform in the number of iterations in \[15,16\]. More recently, it is shown that the convergence is polynomial in terms of dimension and error tolerance \[17,18,20\].
Besides the ULA, higher-order Langevin diffusion for accelerated sampling algorithms are presented in [22, 23]. Analysis of “leapfrog” implementation of Hamiltonian Monte-Carlo (HMC) for strongly log-concave target distributions is presented in [24] and [25]. Following the introduction of a stochastic gradient-based Langevin approach for Bayesian inference in [26], stochastic gradient based Langevin diffusion and other HMC schemes are presented in [27–31].

The approaches discussed so far assume a centralized entity to process large datasets. However, communication challenges associated with transferring large amounts of data to a central location and the associated privacy issues motivate a decentralized implementation over its centralized counterparts [32, 33].

Master-slave architecture for distributed MCMC via moment sharing is presented [34]. Data-parallel MCMC algorithms for large-scale Bayesian posterior sampling are presented in [35–38]. These algorithms are mainly designed for master-slave computing nodes where slave nodes can all communicate with a master node.

Contribution: In this paper, we draw on the recent ULA literature and develop a decentralized learning algorithm based on the centralized ULA in a nonconvex setting. We consider the problem of collaboratively inferring the global posterior of a parameter of interest based on independent data sets distributed among a network of n agents. The communication topology between the agents can be any undirected connected graph, including the master-slave topology as a special case. We propose a decentralized ULA (D-ULA) that incorporates an average consensus process into the ULA with time-varying step-sizes. In this algorithm, each agent shares its current Markov Chain sample with neighboring agents at each time step. We show that the resulting distribution of the averaged sample converges to the true posterior asymptotically. We provide theoretical analysis of the convergence rate and step-size conditions to achieve speed up of convergence with respect to the number of agents. Empirical results show that the performance of our proposed algorithm is on par with centralized ULA with considerable improvement in the convergence rate, for three different machine learning tasks.

Notation: Let $\mathbb{R}^{n \times m}$ denote the set of $n \times m$ real matrices. For a vector $\phi$, $\phi_i$ is the $i$-th entry of $\phi$. An $n \times n$ identity matrix is denoted as $I_n$ and $1_n$ denotes an $n$-dimensional vector of all ones. For $p \in [1, \infty]$, the $p$-norm of a vector $x$ is denoted as $\|x\|_p$. For matrices $A \in \mathbb{R}^{m \times n}$ and $B \in \mathbb{R}^{p \times q}$, $A \otimes B \in \mathbb{R}^{mp \times nq}$ denotes their Kronecker product. For a graph $\mathcal{G}(\mathcal{V}, \mathcal{E})$ of order $n$, $\mathcal{V} \triangleq \{v_1, \ldots, v_n\}$ represents the agents or nodes and the communication links between the agents are represented as $\mathcal{E} \triangleq \{e_1, \ldots, e_\ell\} \subseteq \mathcal{V} \times \mathcal{V}$. Let $\mathcal{A} = [a_{i,j}] \in \mathbb{R}^{n \times n}$ be the adjacency matrix with entries of $a_{i,j} = 1$ if $(v_i, v_j) \in \mathcal{E}$ and zero otherwise. Define $\Delta = \text{diag}(\mathcal{A}1_n)$ as the in-degree matrix and $\mathcal{L} = \Delta - \mathcal{A}$ as the graph Laplacian.

2 Problem formulation

Consider a connected network of $n$ agents, each with a randomly distributed set of $m_i$ data items, $X_i = \{x^i_j\}_{j=1}^{m_i}, \forall i = 1, \ldots, n$. Here $x^i_j \in \mathbb{R}^{d_x}$ is the
The $j$-th data element in a set of $m_i$ data items available to the $i$-th agent. Let $w \in \mathbb{R}^{d_w}$ be the parameter vector associated with the model and $p(w)$ is the prior associated with the model parameters. The global posterior distribution of $w$ given the $n$ independent data sets distributed among the agents can be expressed as

$$p(w|X_1, \ldots, X_n) \propto p(w) \prod_{i=1}^{n} p(X_i|w) = \prod_{i=1}^{n} p(X_i|w)p(w)^{\frac{1}{n}}.$$  

(1)

In the optimization literature, the prior, $p(w)$, regularizes the parameter and the likelihood, $p(X_i|w)$, represents the local cost function available to each agent. Here, the set of $n$ independent data sets are distributed among $n$ agents each with a size of $m_i$, $i = 1, \ldots, n$. At the risk of abusing the notation, define $p(w|X_i)$ as the local posterior distribution. Thus the global posterior can be written as the product of local posteriors as

$$p(w|X_1, \ldots, X_n) \propto \prod_{i=1}^{n} p(w|X_i).$$  

(2)

The main issue with point estimates obtained from optimization schemes like maximum likelihood and maximum a posteriori estimation is that they fail to capture the parameter uncertainty and they can potentially over-fit the data. This paper is aimed at developing a method for collaborative Bayesian learning from large scale datasets distributed among a networked set of agents as a solution to the numerous issues associated with the point estimation schemes. In particular, we present a decentralized version of the unadjusted Langevin algorithm to distributedly obtain samples from the global posterior $p(w|X_1, \ldots, X_n)$. For the ease of notation, we use $X$ to denote the entire data set. Thus the global posterior can be written as $p(w|X)$.

3 Decentralized unadjusted Langevin algorithm

To efficiently explore the global posterior $p(w|X)$, we first rewrite the target distribution in terms of an energy function $U$ as follows [1, 39, 40]:

$$p(w|X) \propto \exp(-U(w)), $$  

(3)

where $U$ is the analogue of potential energy given by

$$U(w) \propto -\log p(w|X).$$  

(4)

The Langevin algorithm is a well known family of gradient based Monte Carlo sampling algorithms. The sample obtained using Unadjusted Langevin Algorithm (ULA) at a given time instant $k$ is given by [19]

$$w(k+1) = w(k) - \alpha_k \nabla U(w(k)) + \sqrt{2\alpha_k}v(k)$$  

(5)

where $\alpha_k$ is the algorithm step-size, $w(k)$ represents the sample obtained at the $k$-th time instant and $v(k)$ is a $d_w$-dimensional, independent, zero-mean, unit
variance, Gaussian sequence, i.e., \(v(k) \sim \mathcal{N}(0, I_{d_w})\), \(\forall k \geq 0\). Now substituting (4) yields
\[
\mathbf{w}(k+1) = \mathbf{w}(k) + \alpha_k \nabla \log p(\mathbf{w}(k)|\mathbf{X}) + \sqrt{2\alpha_k} \mathbf{v}(k).
\] (6)

Substituting (1) yields
\[
\mathbf{w}(k+1) = \mathbf{w}(k) + \alpha_0 \sum_{i=1}^{n} \left( \nabla \log p(\mathbf{X}_i|\mathbf{w}(k)) + \frac{1}{n} \nabla \log p(\mathbf{w}(k)) \right) + \sqrt{2\alpha_k} \mathbf{v}(k).
\] (7)

The samples obtained using the continuous-time version of the centralized ULA given in (4) have shown to exponentially converge to the target posterior distribution (41) for a certain class of distributions with exponential tails. Convergence properties of the ULA had been widely studied for log-concave target distributions [6, 8–12]. Non-asymptotic analysis of centralized ULA without the strong log-concavity assumption on target distribution is presented in [13–20].

However, when the data is distributed among \(n\) agents and there is no central agent to pool all the local gradients, exact implementation of the above ULA is difficult, if not impossible. Therefore we propose the following decentralized ULA:

\[
\mathbf{w}_i(k+1) = \mathbf{w}_i(k) - \beta_k \sum_{j=1}^{n} a_{i,j} (\mathbf{w}_i(k) - \mathbf{w}_j(k))
\]

\[
+ \alpha_k n \left( \nabla \log p(\mathbf{X}_i|\mathbf{w}_i(k)) + \frac{1}{n} \nabla \log p(\mathbf{w}_i(k)) \right) + \sqrt{2\alpha_k} \mathbf{v}_i(k),
\] (8)

where \(a_{i,j}\) denotes the entries of the adjacency matrix corresponding to the communication network \(G(\mathcal{V}, \mathcal{E})\), \(\beta_k\) is the consensus step-size and \(\mathbf{v}_i(k)\) are \(d_w\)-dimensional, independent, zero-mean, unit variance, Gaussian sequence, i.e., \(\mathbf{v}_i(k) \sim \mathcal{N}(0, I_{d_w})\), \(\forall i \in \mathcal{I}\).

Define \(\mathbf{w}(k) \triangleq [\mathbf{w}_1(k) \cdots \mathbf{w}_n(k)]^\top \in \mathbb{R}^{d_w}\) and \(\mathbf{v}(k) \triangleq [\mathbf{v}_1(k) \cdots \mathbf{v}_n(k)]^\top \in \mathbb{R}^{d_w}\). Now (8) can be written as

\[
\mathbf{w}(k+1) = \mathbf{w}(k) - \beta_k (\mathcal{L} \otimes I_{d_w}) \mathbf{w}(k) - \alpha_k n \mathbf{g}(\mathbf{w}(k), \mathbf{X}) + \sqrt{2\alpha_k} \mathbf{v}(k),
\] (9)

where \(\mathcal{L}\) is the network Laplacian and

\[
\mathbf{g}(\mathbf{w}(k), \mathbf{X}) \triangleq \begin{bmatrix}
\mathbf{g}_1 (\mathbf{w}_1(k), \mathbf{X}_1) \\
\vdots \\
\mathbf{g}_n (\mathbf{w}_n(k), \mathbf{X}_n)
\end{bmatrix} = \begin{bmatrix}
\nabla U_1 (\mathbf{w}_1(k), \mathbf{X}_1) \\
\vdots \\
\nabla U_n (\mathbf{w}_n(k), \mathbf{X}_n)
\end{bmatrix}
\]

where \(U_i(\mathbf{w}, \mathbf{X}_i) = -\log p(\mathbf{w}|\mathbf{X}_i)\) and \(p(\mathbf{w}|\mathbf{X}_i)\) is the local posterior, given in (1). Define the network weight-matrix \(\mathcal{W}_k = (I_n - \beta_k \mathcal{L})\). Thus the proposed decentralized ULA can be written as

\[
\mathbf{w}(k+1) = (\mathcal{W}_k \otimes I_{d_w}) \mathbf{w}(k) - \alpha_k n \mathbf{g}(\mathbf{w}(k), \mathbf{X}) + \sqrt{2\alpha_k} \mathbf{v}(k).
\] (10)
If we ignore the additive noise term, then the decentralized ULA of (10) can be considered a consensus optimization algorithm aimed at solving the problem, \( \min_w U(w, X) \), where

\[
U(w, X) = \sum_{i=1}^{n} U_i(w, X_i).
\] (11)

Denote by \( p^* \) the stationary probability distribution corresponding to the global posterior distribution, i.e., the target distribution. It then follows from (3) that \( p^*(\cdot) = \exp(-U(\cdot, X) + C) \), (12) for some positive constant \( C \) corresponding to the normalizing constant. Now note that the centralized ULA for generating samples from the target distribution \( p^*(\bar{w}^*) \) of (12) is given as \( \bar{w}^*(k+1) = \bar{w}^*(k) - \alpha_k \nabla U(\bar{w}^*(k), X) + \sqrt{2\alpha_k} \bar{v}(k) \). (13)

The continuous-time limit of (13) can be obtained as the following Stochastic Differential Equation (SDE) known as the Langevin equation [42]:

\[
d\bar{w}^*(t) = -\nabla U(\bar{w}^*(t), X) dt + \sqrt{2} dB_t,
\] (14)

where \( B_t \) is a \( d_w \)-dimensional Brownian motion. The pseudocode of the proposed decentralized ULA is given in Algorithm 1, where \( \alpha_k = \frac{a}{(k+1)^{\delta_2}} \) and \( \beta_k = \frac{b}{(k+1)^{\delta_1}} \) (see Condition 1 in S1).

4 Main Results

Analysis of the proposed distributed ULA given in (10) requires that the sequences \( \{\alpha_k\} \) and \( \{\beta_k\} \) be selected as (see Condition 1 in S1)

\[
\alpha_k = \frac{a}{(k+1)^{\delta_2}} \quad \text{and} \quad \beta_k = \frac{b}{(k+1)^{\delta_1}},
\] (15)

where \( 0 < a, b, \delta_2 \leq 0, \delta_1 \leq \frac{1}{2}, \delta_1 + \delta_2 < 1 \). Furthermore, we make the following three assumptions (formally stated in S1): (i) the gradients \( \nabla U_i \) are Lipschitz continuous with Lipschitz constant \( L_i > 0 \), \( \forall i = 1, \ldots, n \); (ii) the communication network is given as a connected undirected graph; and (iii) there exists a positive constant \( \mu_g < \infty \) such that the disagreement on the gradient among the distributed agents, denoted as \( g(w, X) \), satisfies \( \mathbb{E} \left[ \|g(w_k, X)\|_2^2 | w_k \right] \leq \mu_g (1 + k)^{\delta_2} \) a.s., where \( g(w_k, X) = g(w_k, X) - (\frac{1}{n} 1_n 1_n^\top \otimes I_{d_w}) g(w_k, X) \).

From the proposed distributed ULA given in (10), the average dynamics is given as

\[
\bar{w}(k+1) = \bar{w}(k) - \alpha_k \sum_{i=1}^{n} \nabla U_i (w_i(k), X_i) + \sqrt{2\alpha_k} \bar{v}(k),
\] (16)
Algorithm 1 Decentralized ULA

1: Initialization : \(w(0) = [w_1^T(0) \ldots w_n^T(0)]^T\)
2: Input : \(a, b, \delta_1\) and \(\delta_2\)
3: for \(k \geq 0\) do
4:   for \(i = 1\) to \(n\) do
5:     Sample \(v_i(k) \sim \mathcal{N}(0, I_{d_{w}})\) & compute \(g_i(w_i(k), X_i)\)
6:     Compute \(\tilde{w}_i(k) = \sum_{j=1}^{n} a_{i,j} (w_i(k) - w_j(k))\)
7:     Update \(w_i(k+1) = w_i(k) - \beta_i \tilde{w}_i(k) - \alpha_i n g_i(w_i(k), \xi_i(k)) + \sqrt{2\alpha_i v_i(k)}\)
8:   end for
9: end for

where \(\tilde{w}(k) = \frac{1}{n} \sum_{i=1}^{n} w_i(k)\) and \(\bar{w}(k) = \frac{1}{n} \sum_{i=1}^{n} v_i(k)\). Since individual \(v_i(k)\) are independent, zero-mean, unit-variance Gaussian random vectors, \(\bar{v}(k)\) is also a zero-mean, unit-variance Gaussian random vector. Now adding and subtracting \(\alpha_k \sum_{i=1}^{n} \nabla U_i(\tilde{w}(k), X_i) = \alpha_k \nabla U(\tilde{w}(k), X)\) yields

\[
\tilde{w}(k+1) = \tilde{w}(k) - \alpha_k \nabla U(\tilde{w}(k), X) - \alpha_k \zeta(\tilde{w}(k), \tilde{w}(k)) + \sqrt{2\alpha_k \bar{v}(k)},
\]

where \(U(\tilde{w}(k), X)\) is defined in (11), the consensus error \(\tilde{w}(k)\) is defined as \(\tilde{w}(k) = ((I_n - \frac{1}{n} 1_n 1_n^T) \otimes I_{d_w}) w(k)\) and \(\zeta(\tilde{w}(k), \tilde{w}(k))\) is defined as

\[
\zeta(\tilde{w}(k), \tilde{w}(k)) = \sum_{i=1}^{n} (\nabla U_i(\tilde{w}(k) + \bar{w}_i(k), X_i) - \nabla U_i(\tilde{w}(k), X_i)).
\]

For all \(k \geq 0\), let \([t_k, t_{k+1}]\) denote the current time-interval, i.e., \(t \in [t_k, t_{k+1}]\), where \(t_k\) is defined as \(t_k = \sum_{j=0}^{k-1} \alpha_j\). Here, \(t_{k+1} = t_k + \alpha_k\). Define \(\tilde{w}(t)\) as

\[
\tilde{w}(t) = \tilde{w}(k), \quad \forall t \in [t_k, t_{k+1}), \quad k \geq 0.
\]

Now (17) can be written as

\[
\tilde{w}(t_{k+1}) = \tilde{w}(t_k) - \alpha_k \nabla U(\tilde{w}(t_k), X) - \alpha_k \zeta(\tilde{w}(t_k), \tilde{w}(t_k)) + \sqrt{2} (B_{t_{k+1}} - B_{t_k}),
\]

where \(B_t\) is a \(d_{w}\)-dimensional Brownian motion. Thus, for \(t_k \leq t < t_{k+1}\), the discretized equation of (17) is given by

\[
d\tilde{w}(t) = -\nabla U(\tilde{w}(t), X) dt + \sqrt{2} dB_t - \zeta(\tilde{w}(t), \tilde{w}(t)) dt.
\]

Let \(\tilde{w}(t)\) in (21) admits a probability distribution \(p_t(\tilde{w})\) for \(t_k \leq t < t_{k+1}\). Here we aim to show that \(p_t(\tilde{w}) \rightarrow p^*\) as \(k \rightarrow \infty\).

4.1 Kullback-Leibler (KL) divergence and log-Sobolev inequality

Sampling can be viewed as optimization in the space of measures, where the objective function in the space of measures attains its minimum at the target distribution. Following [3][4][19][20], we use the relative entropy or the
KL-divergence of \( p_t(\mathbf{w}) \) to the target distribution \( p^* \), denoted by \( F(p_t(\mathbf{w})) \), as the objective, i.e.,
\[
F(p_t(\mathbf{w})) = \int p_t(\mathbf{w}) \log \left( \frac{p_t(\mathbf{w})}{p^*(\mathbf{w})} \right) d\mathbf{w}.
\]  

(22)

KL-divergence is non-negative and it is minimized at the target distribution, i.e., \( F(p_t(\mathbf{w})) \geq 0 \) and \( F(p_t(\mathbf{w})) = 0 \) if and only if \( p_t = p^* \). The property of \( p^* \) that we rely on to show convergence of the proposed algorithm is that it satisfies a log-Sobolev inequality. Consider a Sobolev space defined by the weighted norm: \( \int g(\mathbf{w})^2 p^*(\mathbf{w}) d\mathbf{w} \), where \( p^*(\mathbf{w}) \propto \exp(-U(\mathbf{w})) \). We say that \( p^*(\mathbf{w}) \) satisfies a log-Sobolev inequality if there exists a constant \( \rho_U > 0 \) such that for any smooth function \( g \) satisfying \( \int g(\mathbf{w}) p^*(\mathbf{w}) d\mathbf{w} = 1 \), we have:
\[
\int g(\mathbf{w}) \log g(\mathbf{w}) p^*(\mathbf{w}) d\mathbf{w} \leq \frac{1}{2\rho_U} \int \left\| \nabla g(\mathbf{w}) \right\|^2 p^*(\mathbf{w}) d\mathbf{w},
\]

(23)

where \( \rho_U \) is the log-Sobolev constant. Let \( g(\mathbf{w}) = \frac{p_t(\mathbf{w})}{p^*(\mathbf{w})} \). Thus we have
\[
F(p_t(\mathbf{w})) = \mathbb{E}_{p_t(\mathbf{w})} \left[ \log \left( \frac{p_t(\mathbf{w})}{p^*(\mathbf{w})} \right) \right] \leq \frac{1}{2\rho_U} \mathbb{E}_{p_t(\mathbf{w})} \left[ \left\| \nabla \log \left( \frac{p_t(\mathbf{w})}{p^*(\mathbf{w})} \right) \right\|^2 \right].
\]  

(24)

Now we present our first result, which shows that the average-consensus error \( \tilde{\mathbf{w}}_k \) is decreasing at the rate \( \mathcal{O}\left(\frac{1}{(k+1)^{1.5}}\right) \) (see (S92) for an explicit expression). This implies that the individual samples \( \mathbf{w}_i(k) \) are converging to \( \tilde{\mathbf{w}}_k \) and this is possible only because of the decaying step-size \( \alpha_k \), which also multiplies the additive Gaussian noise.

**Theorem 1.** Consider the decentralized ULA (D-ULA) given in Algorithm 2 under Assumptions 1-3. Then, for the average-consensus error defined as \( \tilde{\mathbf{w}}_k = (I_n - \frac{1}{n} 1_n 1_n^\top \otimes I_{d_n}) \mathbf{w}_k \), there holds:
\[
\mathbb{E} \left[ \left\| \tilde{\mathbf{w}}_{k+1} \right\|^2 \right] \leq \frac{W_3}{\exp(W_1(k+1)^{1-\delta_1})} + \frac{W_2}{(k+1)^{\delta_2-2\delta_1}},
\]  

(25)

where \( W_1, W_2 \) and \( W_3 \) are positive constants defined in (S83), (S84), (S85) and (S86).

Detailed proof of Theorem 1 is given in S3. Now we present our main result which shows that the KL-divergence between \( p_t \) and \( p^* \) is in fact decreasing.

**Theorem 2.** Consider the decentralized ULA (D-ULA) given in Algorithm 2 under Assumptions 1-3 with \( \alpha_k, \beta_k \) given in Condition 1 and \( a \) in \( \alpha_k \) selected as
\[
a = \frac{1}{n^{4/3}} \left( \frac{\rho_U (3\delta_2^2 - 1)}{25L^4 \delta_2^2} \right)^{\frac{1}{2}}.
\]  

(26)
Given that the target distribution satisfies the log-Sobolev inequality \( \rho_U > 0 \) with a constant \( \rho_U > 0 \), and has a bounded second moment, i.e., \( \int \| \bar{w} \|^2_p \cdot \rho^*(\bar{w}) \, d\bar{w} \leq c_1 \) for some bounded positive constant \( c_1 \), then for all initial distributions \( p_{t_0}(\bar{w}) \) satisfying \( F(p_{t_0}(\bar{w})) \leq c_2 \), we have

\[
F(p_{t_{k+1}}(\bar{w})) \leq F(p_{t_0}(\bar{w})) + \bar{C}_{F_1} \exp \left( \rho_U \sum_{\ell=0}^{k} \alpha_\ell \right) + \frac{\bar{C}_{F_2}}{n^{\frac{1}{3}}} \exp \left( \frac{\rho_U}{2} (k+1)^{1-\delta_2} \right) + \frac{\bar{C}_{F_3}}{n^{\frac{1}{3}}} \exp \left( \frac{\rho_U}{2} (k+1)^{1-\delta_2} \right) + \frac{\bar{C}_{F_4}}{n^{\frac{1}{3}}} \exp \left( \frac{\rho_U}{2} (k+1)^{1-\delta_2} \right)
\]

where the positive constants \( \bar{C}_{F_1}, \bar{C}_{F_2}, \bar{C}_{F_3} \) and associated parameters are defined in \((S221) - (S229)\).

Proof of Theorem 2 is given in \( S4 \). Compared to the existing results, by using a decaying step-size, we are able to remove the constant bias term present in the KL-divergence due to the additive noise. In \( (27) \), the constants \( \bar{C}_{F_1} \) and \( \bar{C}_{F_3} \) are dominated by the consensus-error while the additive noise contributes most to \( \bar{C}_{F_2} \). Note that the exponential convergence rate for the initial KL-divergence \( F(p_{t_0}(\bar{w})) \) is similar to what is currently known in the literature \([4, 9, 19]\). More importantly, the constant bias-term present in the existing results for the KL-divergence between the actual and target distribution, which is absorbed into the constant \( \bar{C}_{F_2} \), is decreasing and we do see a speed-up for decay with the number of agents due to the \( n^{1/3} \)-term.

**Corollary 1.** For the decentralized ULA (D-ULA) given in Algorithm 7 under the conditions of Theorem 2 and error tolerance \( \epsilon \in (0, 1) \), there holds

\[
F(p_{t_k}(\bar{w})) \leq \epsilon, \quad \forall k \geq k^* \tag{28}
\]

where \( k^* = \max \left\{ \left( \frac{1}{\rho_U} \log \left( \frac{2Q_1}{\epsilon} \right) \right)^{\frac{1}{1-\delta_2}}, \left( \frac{2Q_2}{\epsilon} \right)^{\frac{1}{1-\delta_2}} \right\} \),

\[
Q_1 = \left( F(p_{t_0}(\bar{w})) + \bar{C}_{F_1} \right) \exp \left( \frac{\rho_U}{1-\delta_2} \right) + \bar{C}_{F_3}
\]

and \( Q_2 = \frac{\bar{C}_{F_2}}{n^{1/3}} \).

Corollary 1 follows from Theorem 2 and the proof is given in \( S5 \). Corollary 1 provides the minimum number of iterations required to decrease the KL-divergence below a given error-tolerance \( \epsilon \).

**5 Numerical experiments**

We apply the proposed algorithm to perform decentralized Bayesian learning for Gaussian mixture modeling, logistic regression, and classification and empirically compare our proposed algorithm to centralized ULA (C-ULA). In all the experiments, we have used a network of five agents in an undirected unweighted ring topology for the decentralized setting. Additional details of all the experiments including step sizes and number of epochs are provided in the Supplementary material (see \( S6 \)).
5.1 Parameter estimation for Gaussian mixture

In this section, we compare the efficiency of D-ULA against the C-ULA for parameter estimation of a multimodal Gaussian mixture with tied means [26].

The Gaussian mixture is given by

\[ \theta_1 \sim \mathcal{N}(0, \sigma_1^2); \quad \theta_2 \sim \mathcal{N}(0, \sigma_2^2) \quad \text{and} \quad x_i \sim \frac{1}{2} \mathcal{N}(\theta_1, \sigma_x^2) + \frac{1}{2} \mathcal{N}(\theta_1 + \theta_2, \sigma_x^2) \]

where \( \sigma_1^2 = 10 \), \( \sigma_2^2 = 1 \), \( \sigma_x^2 = 2 \) and \( \mathbf{w} \triangleq [\theta_1, \theta_2]^T \in \mathbb{R}^2 \). For the centralized setting, similar to [26], 100 data samples are drawn from the model with \( \theta_1 = 0 \) and \( \theta_2 = 1 \). The 100 data samples are randomly divided into 5 sets of 20 samples that are made available to each agent in the decentralized network. The posterior distribution of the parameters is bimodal with negatively correlated modes at \( \mathbf{w} = [0, 1] \) and \( \mathbf{w} = [1, -1] \). As shown in Figure 1, the posteriors estimated by D-ULA and C-ULA replicate the true posterior distribution of parameters. Besides, the posterior plots indicate D-ULA better replicate the true posterior with samples concentrated near the modes compared to C-ULA.

5.2 Bayesian logistic regression

We compare the performance of D-ULA and C-ULA for Bayesian inference of logistic regression models using a9a dataset available at the UCI machine learning repository[1]. The dataset contains 32561 observations and 123 parameters. We use a Laplace prior with a scale of 1 on the parameters. Test accuracy averaged over 50 runs for both approaches are shown in Figure 2. Similar to [26], during each run, we randomly choose 80% of data for training and remaining 20% for testing for the centralized setting. For D-ULA, the training data is randomly divided into 5 sets, each of which is made available to each agent in the decentralized network. During each run, the same 20% partition is used to

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[1] http://www.csie.ntu.edu.tw/~cjlin/libsvmtools/datasets/binary/a9a
test the performance of C-ULA and D-ULA. Test results indicate that the performance of D-ULA with accuracy of 84.38% in 1040 iterations is comparable to that of C-ULA which settles to an accuracy of 83.89% in 5210 iterations. From the inset in Figure 2a, it can be seen that D-ULA converges much faster than C-ULA. Test accuracy of all the agents in D-ULA settles to the same level as shown in the inset of Figure 2b. The shaded region in the figures indicates 1 standard deviation.

5.3 Bayesian learning for handwritten digit classification and OOD detection

In this section, we present decentralized Bayesian learning as a potential strategy to recognize handwritten digits in images. For this, we use the MNIST data set containing 60000 gray scale images of 10 digits (0-9) for training and 10000 images for testing. Each agent in D-ULA aims to train its own neural network, which is a randomly initialized LeNet-5 with Kaiming uniform prior on the parameters of the network. Each agent has access to 12000 randomly chosen training samples. Test accuracy obtained using stochastic gradient descent (SGD), C-ULA, and 5 agents of D-ULA after 10 epochs are 98.15%, 98.16%, 98.52%, 98.52%, 98.39%, 98.45% and 98.47%, respectively.

Next, we explore the efficacy of the proposed algorithm to detect out-of-distribution (OOD) samples or outliers in the datasets. We train each LeNet-5 neural network on the MNIST training data set and test it on MNIST test data set for normalcy class and Street View House Numbers (SVHN) test data set for OOD data. SVHN data set is similar to MNIST, but with color images of 10 digits (0-9) and extra confusing digits around the central digit of interest. We converted them to gray scale for this experiment. Networks trained on MNIST are expected to give relatively low prediction probabilities for SVHN data samples. Table II summarizes the mean and standard deviation of probabilities of predicted labels obtained for all the approaches. Since SGD is a maximum a posteriori point estimate, it fails to recognize out of sample data sets, and gives high prediction probabilities even for OOD SVHN data.

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2http://ufldl.stanford.edu/housenumbers/
Table 1: Probability of predicted labels (mean/standard deviation)

|          | SGD   | C-ULA | Agent 1 | Agent 2 | Agent 3 | Agent 4 | Agent 5 |
|----------|-------|-------|---------|---------|---------|---------|---------|
| **MNIST** |       |       |         |         |         |         |         |
| Mean     | 0.974 | 0.968 | 0.973   | 0.972   | 0.972   | 0.973   | 0.973   |
| Std. dev. | 0.078 | 0.086 | 0.079   | 0.08    | 0.08    | 0.08    | 0.08    |
| **SVHN** |       |       |         |         |         |         |         |
| Mean     | 0.849 | 0.604 | 0.659   | 0.6588  | 0.653   | 0.663   | 0.651   |
| Std. dev. | 0.154 | 0.169 | 0.188   | 0.189   | 0.188   | 0.19    | 0.187   |

One the other hand, C-ULA and D-ULA show an improved performance in detecting OOD SVHN data by giving lower prediction probabilities for SVHN data, but giving high prediction probabilities for MNIST test data as seen in Table 1. The plots of probability density of predicted labels corresponding to all the approaches are provided in the Supplementary material.

The decentralized ULA results in Section 5.2 and 5.3 were obtained using a “mini-batch” version of the proposed D-ULA algorithm, where the log-likelihood was obtained from random mini-batches of $X_i$ for agent $i$, $i = 1, \ldots, n$. Although our theoretical analysis is based on the likelihood from the entire $X_i$, the empirical results in these two sections show that the “mini-batch” D-ULA algorithm is also effective. This is plausible since the additive noise $\sqrt{2\alpha_k}v_i(k)$ in (9) will dominate the noise in the local posterior term as $k$ increases.

6 Conclusion

In this paper, we present a decentralized collaborative approach for a group of agents to sample the posterior distribution of a parameter of interest with locally available data sets. We assume an undirected connected communication topology between the agents. We propose a decentralized unadjusted Langevin algorithm with time-varying step-sizes and establish conditions on the step-sizes for asymptotic convergence to the target distribution. The algorithm also exhibits a guaranteed speed-up in convergence in the number of agents. We conducted three experiments on Gaussian mixtures, logistic regression, and image classification. The experimental results demonstrated that the proposed algorithm offers improved accuracy with enhanced speed of convergence. The results from the last experiment also suggest a potential application of the proposed algorithm for outlier detection.

Broader Impact

This work presents a basic line of research on reducing computational complexity, enhancing speed of convergence, and addressing potential privacy issues associated with centralized Bayesian learning. Experiments and empirical results cover a broad set of applications including parameter estimation for local non-convex models, logistic regression, image classification and outlier detection. We have used publicly available datasets, which have no implications on machine learning bias, fairness or ethics. Hence, we believe that this section about potential negative impact of our work on society is not applicable to the proposed work.
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Decentralized Langevin Dynamics for Bayesian Learning

(Supplementary Material)

S1 Assumptions

We first make the following assumption regarding the $U_i$:

**Assumption 1.** The gradients $\nabla U_i$ are Lipschitz continuous with Lipschitz constant $L_i > 0$, $\forall i = 1, \ldots, n$, i.e., $\forall w_a, w_b \in \mathbb{R}^{d_w}$

$$\|\nabla U_i(w_a, X_i) - \nabla U_i(w_b, X_i)\|_2 \leq L_i\|w_a - w_b\|_2.$$  \hspace{1cm} (S1)

Let

$$U(w, X) = \sum_{i=1}^{n} U_i(w, X_i).$$ \hspace{1cm} (S2)

Following Assumption 1, the function $U$ is continuously differentiable and the gradient $\nabla U$ is Lipschitz continuous, i.e., $\forall w_a, w_b \in \mathbb{R}^{d_w}$, there exists a positive constant $\bar{L}$ such that

$$\|\nabla U(w_a, X) - \nabla U(w_b, X)\|_2 \leq \bar{L}\|w_a - w_b\|_2,$$ \hspace{1cm} (S3)

where $L = \max \{L_i\}$ and $\nabla F(w, X) \triangleq [\nabla U_1(w_1, X_1) \ldots \nabla U_n(w_n, X_n) \top \top] \top \in \mathbb{R}^{nd_w}$.

**Assumption 2.** The interaction topology of $n$ networked agents is given as a connected undirected graph $G(V, E)$.

For the connected undirected graph $G(V, E)$, the graph Laplacian $L$ is a positive semi-definite matrix with one eigenvalue at 0 corresponding to the eigenvector $1_n$. Furthermore, it follows from Lemma 3 [45] that for all $x \in \mathbb{R}^n$, such that $1_n^\top x = 0$, we have $x^\top L x = 0$.

Here we require the following condition on $\{\alpha_k\}$ and $\{\beta_k\}$:

$\hspace{12cm}$\footnote{Unless otherwise specified, gradients are taken with respect to the parameter $w$ or its aggregate $w$.}
Condition 1. Sequences \( \{\alpha_k\} \) and \( \{\beta_k\} \) are selected as

\[
\alpha_k = \frac{a}{(k+1)^{\delta_2}} \quad \text{and} \quad \beta_k = \frac{b}{(k+1)^{\delta_1}}.
\]

where \( 0 < a, 0 < b, 0 \leq \delta_1 \) and \( \frac{1}{2} + \delta_1 < \delta_2 < 1 \). Also, the parameter \( b = \beta_0 \) in sequence \( \{\beta_k\} \) is selected such that \( \mathcal{W}_0 = (I_n - b\mathcal{L}) \) has a single eigenvalue at 1 corresponding to the right and left eigenvectors \( \mathbf{1}_n \) and \( \mathbf{1}_n^\top \), respectively. Furthermore, the remaining \( n-1 \) eigenvalues of \( \mathcal{W}_0 \) are strictly inside the unit circle.

For sequences \( \{\alpha_k\} \) and \( \{\beta_k\} \) that satisfy Condition 1 we have \( \sum_{k=0}^{\infty} \alpha_k = \infty, \sum_{k=0}^{\infty} \beta_k^2 = \infty \) and \( \sum_{k=0}^{\infty} \alpha_k^2 \beta_k^2 = \infty \). Thus \( \alpha_k, \beta_k, \beta^2_k \) and \( \beta^2_k \) are not summable sequences while \( \alpha_k \) is square-summable. Also note that \( \beta_k \) is allowed to be a constant for all \( k \geq 0 \). However, \( b \) is selected such that \( b < 1/\sigma_{\max}(\mathcal{L}) \), where \( \sigma_{\max}(\cdot) \) denotes the largest singular value. Thus, \( b\sigma_{\max}(\mathcal{L}) < 1 \).

Let \( \mathcal{F}_k \) denotes a filtration generated by the sequence \( \{w_0, \ldots, w_k\} \), i.e.,

\[
\mathbb{E}[v_k | \mathcal{F}_k] = 0
\]

\[
\mathbb{E}[w_{k+1} | \mathcal{F}_k] = (\mathcal{W}_k \otimes I_{d_w}) w_k - \alpha_k n \mathbb{E}[g(w_k, X) | \mathcal{F}_k] \quad \text{a.s.,}
\]

where a.s. (almost surely) denotes events that occur with probability one. Let

\[
\hat{g}(w_k, X) = g(w_k, X) - \left( \frac{1}{n} \mathbf{1}_n \mathbf{1}_n^\top \otimes I_{d_w} \right) g(w_k, X).
\]

Note that \( \hat{g}(w_k, X) \) denotes the disagreement on the gradient among the distributed agents. Here we make the following assumption regarding \( \hat{g}(w_k, X) \):

Assumption 3. There exists a positive constant \( \mu_g < \infty \) such that

\[
\mathbb{E} \left[ \|\hat{g}(w_k, X)\|_2^2 | \mathcal{F}_k \right] \leq \mu_g (1 + k)^{\delta_2} \quad \text{a.s.,}
\]

where \( \delta_2 \) is defined in Condition 1.

Note that Assumption 3 does not uniformly bound \( \mathbb{E} \left[ ||\hat{g}(w_k, X)||_2^2 | \mathcal{F}_k \right] \). In fact \( \mathbb{E} \left[ ||\hat{g}(w_k, X)||_2^2 | \mathcal{F}_k \right] \) can grow unbounded with time, i.e., as \( k \rightarrow \infty \).

S2 Useful Lemmas

Lemma S1. Given Assumption 3, for \( U \) defined in (11), we have \( \forall w_a, w_b \in \mathbb{R}^{d_w}, \)

\[
U(w_b, X) \leq U(w_a, X) + \nabla U(w_a, X)^\top (w_b - w_a) + \frac{1}{2} n L \| w_b - w_a \|_2^2.
\]

(S10)

Proof: Proof follows from the mean value theorem.

Lemma S2. Given Assumption 4 we have

\[
M \triangleq \left( I_n - \frac{1}{n} \mathbf{1}_n \mathbf{1}_n^\top \right) = \mathcal{L}(\mathcal{L}),
\]

(S11)
where $(\cdot)^+$ denotes the generalized inverse. Furthermore, for all $x \in \mathbb{R}^n$ such that $x \notin \mathbb{R}_1^n$, we have
\[
\tilde{x}^T \mathcal{L} \tilde{x} = x^T \mathcal{L} x > \lambda_2(\mathcal{L}) x^T x,
\]
(S12)
where $\tilde{x} = Mx$ is the average-consensus error and $\lambda_2(\mathcal{L})$ denotes the second smallest eigenvalue of $\mathcal{L}$.

**Proof:** For the connected undirected graph, $\mathcal{L}$ is a positive semi-definite matrix with one eigenvalue at 0 corresponding to the eigenvector $1_n$. Thus
\[
\mathcal{L} \tilde{x} = \mathcal{L} \left( I_n - \frac{1}{n} 1_n 1_n^T \right) x = x^T \mathcal{L} x,
\]
and for $x \in \mathbb{R}^n$ such that $x \notin \mathbb{R}_1^n$, we have $x^T \mathcal{L} x > \lambda_2(\mathcal{L}) x^T x$. See Lemma 3 of [45] for a detailed proof of (S11).

**Lemma S3.** Let $f(k)$ be a non-negative and decreasing sequence for all $k \geq k_0$. Then for all $k \leq K$, we have
\[
\int_k^K f(x) \, dx \leq \sum_{t=k}^K f(t) \leq \int_{k-1}^K f(x) \, dx.
\]
(S13)
Furthermore, if $f(k)$ is non-negative and increasing, then for all $k \leq K$ we have
\[
\int_{k-1}^K f(x) \, dx \leq \sum_{t=k}^K f(t) \leq \int_k^{K+1} f(x) \, dx.
\]
(S14)

**Proof:** See Appendix A2 in [46].

**Lemma S4.** For all $k \geq 0$, let $y_k$ be a nonnegative sequence satisfying:
\[
y_{k+1} \leq \left( 1 - \frac{\mu_\beta}{(k+1)^{\delta_1}} \right) y_k + \frac{\mu_\zeta}{(k+1)^{\delta_4}},
\]
(S15)
where $0 < \mu_\beta \leq 1$, $0 < \mu_\zeta$, $0 \leq \delta_1 < 1$ and $\delta_1 < \delta_4$ are positive constants. Then we have
\[
y_{k+1} \leq \frac{Y_3}{\exp(Y_1 (k+1)^{1-\delta_1})} + \frac{Y_2}{(k+1)^{\delta_4-\delta_1}},
\]
(S16)
where the constants $Y_1$, $Y_2$ and $Y_3$ are defined as
\[
Y_1 = \frac{\mu_\beta}{1-\delta_1},
\]
(S17)
\[
Y_2 = \frac{\mu_\zeta \delta_4}{\mu_\beta \delta_1} \exp(Y_1 2^{1-\delta_1}),
\]
(S18)
\[
Y_3 = \exp(Y_1) \left( y_0 + \sum_{t=0}^k \left( \frac{1}{(1-\mu_\beta)^t (t+1)^{\delta_4}} \right) \right),
\]
(S19)
where $y_0$ is the initial condition and $\kappa > 0$ is defined as

$$\kappa = \left[ \left( \frac{\delta_1}{\mu_\beta} \right)^{1-\delta_1} \right]. \quad (S20)$$

**Proof:** Let

$$\beta_k = \frac{\mu_\beta}{(k+1)^{\delta_1}}, \quad (S21)$$

$$\zeta_k = \frac{\mu_\zeta}{(k+1)^{\delta_1}}, \quad (S22)$$

and

$$\eta_k = (1 - \beta_k). \quad (S23)$$

Now (S15) can be written as

$$y_{k+1} \leq \eta_k y_k + \zeta_k = \zeta_k + y_0 \prod_{t=0}^{k} \eta_t + \sum_{t=0}^{k-1} \zeta_t \left( \prod_{i=t+1}^{k} \eta_i \right) \quad (S24)$$

Since empty product is 1, we have

$$y_{k+1} \leq y_0 \prod_{t=0}^{k} \eta_t + \sum_{t=0}^{k} \zeta_t \left( \prod_{i=t+1}^{k} \eta_i \right) \quad (S25)$$

Note that $\eta_k \leq 1$ and $\eta_k \to 1$ as $k \to \infty$. Then, using

$$1 - \varphi \leq \exp(-\varphi), \quad 0 \leq \varphi \leq 1 \quad (S26)$$

yields

$$\prod_{t=0}^{k} \eta_t = \prod_{t=0}^{k} (1 - \beta_t) \leq \exp \left( - \sum_{t=0}^{k} \beta_t \right) \quad (S27)$$

Since $\beta_k$ is monotonically decreasing, from Lemma S3 we have

$$\sum_{t=0}^{k} \beta_t \geq \int_{0}^{k} \frac{\mu_\beta}{(t+1)^{\delta_1}} \, dt = \frac{\mu_\beta (k+1)^{1-\delta_1}}{1 - \delta_1} - \frac{\mu_\beta}{1 - \delta_1} \quad (S28)$$

Thus

$$\prod_{t=0}^{k} \eta_t \leq \exp \left( - \sum_{t=0}^{k} \beta_t \right) \leq \frac{\exp(Y_1)}{\exp(Y_1 (k+1)^{1-\delta_1})}, \quad (S29)$$

where

$$Y_1 = \frac{\mu_\beta}{1 - \delta_1}.$$
Similarly
\[ \prod_{i=t+1}^{k} \eta_i = \prod_{i=t+1}^{k} (1 - \beta_i) \leq \exp \left( - \sum_{i=t+1}^{k} \beta_i \right) \]  
(S30)

and
\[ \sum_{i=t+1}^{k} \beta_i \geq \int_{t+1}^{k} \frac{\mu\beta}{(x+1)^{\delta_1}} \, dx = \frac{\mu\beta(k+1)^{1-\delta_1}}{1-\delta_1} - \frac{\mu\beta(t+2)^{1-\delta_1}}{1-\delta_1} \]  
(S31)

Thus
\[ \exp \left( - \sum_{i=t+1}^{k} \beta_i \right) \leq \exp \left( - \frac{\mu\beta(k+1)^{1-\delta_1}}{(1-\delta_1)} + \frac{\mu\beta(t+2)^{1-\delta_1}}{(1-\delta_1)} \right), \]  
(S32)

\[ = \exp \left( - Y_1(k+1)^{1-\delta_1} + Y_1(t+2)^{1-\delta_1} \right). \]  
(S33)

Note that for some \( \bar{t} \in (0, k) \), we have
\[ \sum_{t=0}^{k} \zeta_t \left( \prod_{i=t+1}^{k} \eta_i \right) = \sum_{t=0}^{\bar{t}} \zeta_t \left( \prod_{i=t+1}^{k} \eta_i \right) + \sum_{t=\bar{t}+1}^{k} \zeta_t \left( \prod_{i=t+1}^{k} \eta_i \right) \]  
(S34)

and
\[ \sum_{t=\bar{t}+1}^{k} \zeta_t \left( \prod_{i=t+1}^{k} \eta_i \right) \leq \sum_{t=\bar{t}+1}^{k} \mu_\zeta \exp \left( - Y_1(k+1)^{1-\delta_1} + Y_1(t+2)^{1-\delta_1} \right) \]  
(S35)

\[ = \mu_\zeta \exp \left( - Y_1(k+1)^{1-\delta_1} \right) \sum_{t=\bar{t}+1}^{k} \frac{\exp \left( Y_1(t+2)^{1-\delta_1} \right)}{(t+1)^{\delta_4}} \]  
(S36)

\[ \leq \mu_\zeta \exp \left( - Y_1(k+1)^{1-\delta_1} \right) \sum_{t=\bar{t}+1}^{k} \frac{\exp \left( Y_1t^{1-\delta_1} + Y_12^{1-\delta_1} \right)}{(t+1)^{\delta_4}} \]  
(S37)

\[ \leq \mu_\zeta \exp \left( - Y_1(k+1)^{1-\delta_1} \right) \sum_{t=\bar{t}+1}^{k} \frac{\exp \left( Y_1t^{1-\delta_1} \right)}{t^{\delta_4}} \]  
(S38)

Now it follows from Lemma S3 that
\[ \sum_{t=\bar{t}+1}^{k} \frac{\exp \left( Y_1t^{1-\delta_1} \right)}{t^{\delta_4}} \leq \int_{\bar{t}}^{k+1} \frac{\exp \left( Y_1t^{1-\delta_1} \right)}{t^{\delta_4}} \, dt. \]  
(S39)
Thus we have
\[ \sum_{t=t+1}^k \zeta_t \left( \prod_{i=t+1}^k \eta_i \right) \leq \mu_\zeta \frac{\exp(Y_1^{21-\delta_i})}{\exp(Y_1(k+1)^{1-\delta_i})} \int_{\tilde{t}}^{k+1} \frac{\exp(Y_1 t^{1-\delta_i})}{t^{\delta_4}} \, dt \]  
(S40)

We note
\[ \frac{d}{dt} (\exp(Y_1 t^{1-\delta_i}) t^{-\delta_4+\delta_1}) = Y_1(1 - \delta_1) \exp(Y_1 t^{1-\delta_i}) t^{-\delta_4} - (\delta_4 - \delta_1) \exp(Y_1 t^{1-\delta_i}) t^{-\delta_4+\delta_1-1} \]  
(S41)

Thus for
\[ t \geq \tilde{t} = \left( \frac{\delta_4}{Y_1(1 - \delta_1)} \right) \frac{1}{t^{\delta_4}} = \left( \frac{\delta_4}{\mu_\beta} \right) \frac{1}{t^{\delta_4}} \]  
(S43)

we have
\[ \left( Y_1(1 - \delta_1) - \frac{(\delta_4 - \delta_1)}{t^{1-\delta_1}} \right) \geq \frac{\mu_\beta \delta_1}{\delta_4} \]  
(S44)

and
\[ \frac{d}{dt} (\exp(Y_1 t^{1-\delta_i}) t^{-\delta_4+\delta_1}) \geq \frac{\mu_\beta \delta_1}{\delta_4} \exp(Y_1 t^{1-\delta_i}) t^{-\delta_4}. \]  
(S45)

Thus we have
\[ \frac{\exp(Y_1 t^{1-\delta_i})}{t^{\delta_4}} \leq \frac{\delta_4}{\mu_\beta \delta_1} \frac{d}{dt} (\exp(Y_1 t^{1-\delta_i}) t^{-\delta_4+\delta_1}) \]  
(S46)

and
\[ \int_{\tilde{t}}^{k+1} \frac{\exp(Y_1 t^{1-\delta_i})}{t^{\delta_4}} \, dt \leq \frac{\delta_4}{\mu_\beta \delta_1} \left( \frac{\exp(Y_1 t^{1-\delta_i})}{t^{\delta_4-\delta_1}} \right) \left( \frac{\exp(Y_1(k+1)^{1-\delta_i})}{(k+1)^{\delta_4-\delta_1}} - \frac{\exp(Y_1(\tilde{t})^{1-\delta_i})}{(\tilde{t})^{\delta_4-\delta_1}} \right) \]  
(S47)

Therefore we have
\[ \sum_{t=t+1}^k \zeta_t \left( \prod_{i=t+1}^k \eta_i \right) \leq \frac{\mu_\zeta \exp(Y_1 2^{1-\delta_i})}{\exp(Y_1(k+1)^{1-\delta_i})} \frac{\delta_4}{\mu_\beta \delta_1} \left( \frac{\exp(Y_1(k+1)^{1-\delta_i})}{(k+1)^{\delta_4-\delta_1}} - \frac{\exp(Y_1(\tilde{t})^{1-\delta_i})}{(\tilde{t})^{\delta_4-\delta_1}} \right) \]  
(S49)

\[ = \frac{\mu_\zeta \exp(Y_1 2^{1-\delta_i})}{\exp(Y_1(k+1)^{1-\delta_i})} \left( \frac{\delta_4 \exp(Y_1(k+1)^{1-\delta_i})}{\mu_\beta \delta_1 (k+1)^{\delta_4-\delta_1}} - Y_4 \right) \]  
(S50)
where $Y_4$ is a positive constant defined as

$$Y_4 = \frac{\delta_4 \exp \left( Y_1 (t)^{1-\delta_1} \right)}{\mu_\beta \delta_1 (t)^{\delta_4 - \delta_1}}$$  \hspace{1cm} (S51)

Therefore

$$\sum_{t=t+1}^k \zeta_t \left( \prod_{i=t+1}^k \eta_i \right) \leq \frac{\mu_\zeta \exp \left( Y_1 2^{1-\delta_1} \right)}{\exp \left( Y_1 (k + 1)^{1-\delta_1} \right)} \left( \frac{\delta_4 \exp \left( Y_1 (k + 1)^{1-\delta_1} \right)}{\mu_\beta \delta_1 (k + 1)^{\delta_4 - \delta_1}} - Y_4 \right)$$  \hspace{1cm} (S52)

$$= \left( \frac{\mu_\zeta \delta_4 \exp \left( Y_1 2^{1-\delta_1} \right)}{\mu_\beta \delta_1 (k + 1)^{\delta_4 - \delta_1}} - Y_4 \mu_\zeta \exp \left( Y_1 2^{1-\delta_1} \right) \right) \left( \frac{\delta_4 \exp \left( Y_1 (k + 1)^{1-\delta_1} \right)}{\mu_\beta \delta_1 (k + 1)^{\delta_4 - \delta_1}} \right)$$  \hspace{1cm} (S53)

Thus we have

$$\sum_{t=t+1}^k \zeta_t \left( \prod_{i=t+1}^k \eta_i \right) \leq \frac{\mu_\zeta \delta_4 \exp \left( Y_1 2^{1-\delta_1} \right)}{\mu_\beta \delta_1 (k + 1)^{\delta_4 - \delta_1}}$$  \hspace{1cm} (S54)

Now going back to (S25), we can write

$$y_{k+1} \leq y_0 \prod_{t=0}^k \eta_t + \sum_{t=0}^k \zeta_t \left( \prod_{i=t+1}^k \eta_i \right)$$  \hspace{1cm} (S55)

$$= y_0 \prod_{t=0}^k \eta_t + \sum_{t=0}^i \zeta_t \left( \prod_{i=t+1}^k \eta_i \right) + \sum_{t=t+1}^k \zeta_t \left( \prod_{i=t+1}^k \eta_i \right)$$  \hspace{1cm} (S56)

$$= y_0 \prod_{t=0}^k \eta_t + \sum_{t=0}^i \zeta_t \frac{\prod_{i=0}^k \eta_i}{\prod_{i=0}^t \eta_i} + \sum_{t=t+1}^k \zeta_t \left( \prod_{i=t+1}^k \eta_i \right)$$  \hspace{1cm} (S57)

$$= \left( y_0 + \sum_{t=0}^i \frac{\zeta_t}{\prod_{i=0}^t \eta_i} \right) \prod_{t=0}^k \eta_t + \sum_{t=t+1}^k \zeta_t \left( \prod_{i=t+1}^k \eta_i \right).$$  \hspace{1cm} (S58)

Note that since $\eta_k \leq 1$ and $\eta_k \to 1$ as $k \to \infty$, we have

$$\prod_{i=0}^t \eta_i \geq \prod_{i=0}^t \eta_0 = (1 - \mu_\beta)^t$$  \hspace{1cm} (S59)

Thus

$$y_0 + \sum_{t=0}^i \frac{\zeta_t}{\prod_{i=0}^t \eta_i} \leq y_0 + \sum_{t=0}^i \frac{1}{(1 - \mu_\beta)^t (t + 1)^{\delta_4}}$$  \hspace{1cm} (S60)

Now define a bounded constant

$$Y_5 \triangleq \left( y_0 + \sum_{t=0}^i \frac{1}{(1 - \mu_\beta)^t (t + 1)^{\delta_4}} \right)$$  \hspace{1cm} (S61)
Thus we have

\[
y_{k+1} \leq \frac{\exp(Y_1)Y_5}{\exp(Y_1(k+1)^{1-\delta_1})} + \frac{\mu \delta_4 \exp(Y_12^{1-\delta_1})}{\mu \delta_1 (k+1)^{\delta_1 - \delta_1}}
\]  

(S62)

Now (S10) follows from noting that \( Y_3 = \exp(Y_1)Y_5 \) and substituting for \( Y_2 \). ■

S3 Proof of Theorem 1

Consider the DULA given in (10)

\[
w_{k+1} = (W_k \otimes I_{d_w}) w_k - \alpha_k n g(w_k, X) + \sqrt{2} \alpha_k v_k. \tag{S63}
\]

Define the average-consensus error as \( \bar{w}_k = (M \otimes I_{d_w}) w_k \), where \( M = I_n - \frac{1}{n} 1_n 1_n^T \). Thus we have

\[
\bar{w}_{k+1} = (W_k \otimes I_{d_w}) \bar{w}_k - \alpha_k n \bar{g}(w_k, X) + \sqrt{2} \alpha_k \bar{v}_k \tag{S64}
\]

where \( \bar{g}(w_k, X) = (M \otimes I_{d_w}) g(w_k, X) \), \( \bar{v}_k = (M \otimes I_{d_w}) v_k \) and we used the identities \( M (I_n - \beta_k \mathcal{L}) = M - \beta_k \mathcal{L} \) and \( (\mathcal{L} \otimes I_{d_w}) w_k = (\mathcal{L} \otimes I_{d_w}) \bar{w}_k \). Taking the norm on both sides yields

\[
\|\bar{w}_{k+1}\|_2 \leq \|((I_n - \beta_k \mathcal{L}) \otimes I_{d_w}) \bar{w}_k\|_2 + \alpha_k n \|\bar{g}(w_k, X)\|_2 + \sqrt{2} \alpha_k \|\bar{v}_k\|_2. \tag{S65}
\]

Since \( 1_{nd_w}^T \bar{w}_k = 0 \), it follows from [17, Lemma 4.4] that

\[
\|(I_n - \beta_k \mathcal{L}) \otimes I_{d_w}) \bar{w}_k\|_2 \leq (1 - \beta_k \lambda_2(\mathcal{L})) \|\bar{w}_k\|_2, \tag{S66}
\]

where \( \lambda_2(\cdot) \) denotes the second smallest eigenvalue. Thus we have

\[
\|\bar{w}_{k+1}\|_2 \leq (1 - \beta_k \lambda_2(\mathcal{L})) \|\bar{w}_k\|_2 + \sqrt{2} \alpha_k \|\bar{v}_k\|_2 + \alpha_k n \|\bar{g}(w_k, X)\|_2. \tag{S67}
\]

Now we use the following inequality

\[
(x + y)^2 \leq (1 + \theta) x^2 + \left(1 + \frac{1}{\theta}\right) y^2, \tag{S68}
\]

for all \( x, y \in \mathbb{R} \) and \( \theta > 0 \). Since \( \beta_k \lambda_2(\mathcal{L}) < 1 \) for all \( k \geq 0 \), selecting

\[
\theta = (1 - \beta_k \lambda_2(\mathcal{L}))^{-\frac{1}{2}} - 1
\]

yields

\[
\|\bar{w}_{k+1}\|_2^2 \leq (1 - \beta_k \lambda_2(\mathcal{L}))^{-\frac{1}{2}} \left( (1 - \beta_k \lambda_2(\mathcal{L})) \|\bar{w}_k\|_2 + \sqrt{2} \alpha_k \|\bar{v}_k\|_2 \right)^2
\]

\[
+ n^2 \alpha_k^2 \left( \frac{(1 - \beta_k \lambda_2(\mathcal{L}))^{-\frac{1}{2}}}{(1 - \beta_k \lambda_2(\mathcal{L}))^{-\frac{1}{2}} - 1} \right) \|\bar{g}(w_k, X)\|_2^2 \tag{S69}
\]

25
\[(1 - \beta_k \lambda_2(L))^{-\frac{1}{2}} \left( (1 - \beta_k \lambda_2(L)) \|\tilde{w}_k\|_2 + \sqrt{2\alpha_k} \|\tilde{v}_k\|_2 \right)^2 + n^2 \alpha_k^2 \left( \frac{1}{1 - (1 - \beta_k \lambda_2(L))^{\frac{1}{2}}} \right)^2 \|\tilde{g}(w_k, X)\|_2^2 \]

(S70)

Since \(\beta_k \lambda_2(L) < 1\) for all \(k \geq 0\), we have

\[ (1 - \beta_k \lambda_2(L))^{-\frac{1}{2}} \leq \left( 1 - \frac{\beta_k \lambda_2(L)}{2} \right), \]

which results in

\[ \left( \frac{1}{1 - (1 - \beta_k \lambda_2(L))^{\frac{1}{2}}} \right) \leq \left( \frac{1}{1 - \left( 1 - \frac{\beta_k \lambda_2(L)}{2} \right)} \right) = \frac{2}{\beta_k \lambda_2(L)} \]

(S72)

Now it follows from (S70) that

\[
\|\tilde{w}_{k+1}\|_2^2 \leq (1 - \beta_k \lambda_2(L))^{-\frac{1}{2}} \left( (1 - \beta_k \lambda_2(L)) \|\tilde{w}_k\|_2 + \sqrt{2\alpha_k} \|\tilde{v}_k\|_2 \right)^2 + \left( \frac{2n^2 \alpha_k^2}{\beta_k \lambda_2(L)} \right) \|\tilde{g}(w_k, X)\|_2^2
\]

(S73)

Again applying (S68) with the same \(\theta\) yields

\[
\left( (1 - \beta_k \lambda_2(L)) \|\tilde{w}_k\|_2 + \sqrt{2\alpha_k} \|\tilde{v}_k\|_2 \right)^2 \leq (1 - \beta_k \lambda_2(L))^{-\frac{1}{2}} \left( (1 - \beta_k \lambda_2(L))^2 \|\tilde{w}_k\|_2^2 + \left( \frac{(1 - \beta_k \lambda_2(L))^{-\frac{1}{2}}}{(1 - \beta_k \lambda_2(L))^{-\frac{1}{2}} - 1} \right) 2\alpha_k \|\tilde{v}_k\|_2^2 \right)
\]

(S74)

\[
= (1 - \beta_k \lambda_2(L))^{-\frac{1}{2}} \|\tilde{w}_k\|_2^2 + \left( \frac{(1 - \beta_k \lambda_2(L))^{-\frac{1}{2}}}{(1 - \beta_k \lambda_2(L))^{-\frac{1}{2}} - 1} \right) 2\alpha_k \|\tilde{v}_k\|_2^2
\]

(S75)

\[
\leq (1 - \beta_k \lambda_2(L))^{-\frac{1}{2}} \|\tilde{w}_k\|_2^2 + \left( \frac{4\alpha_k}{\beta_k \lambda_2(L)} \right) \|\tilde{v}_k\|_2^2
\]

(S76)

Combining (S73) and (S76) yields

\[
\|\tilde{w}_{k+1}\|_2^2 \leq (1 - \beta_k \lambda_2(L)) \|\tilde{w}_k\|_2^2 + \left( \frac{4\alpha_k}{\beta_k \lambda_2(L)} \right) \|\tilde{v}_k\|_2^2 + \left( \frac{2n^2 \alpha_k^2}{\beta_k \lambda_2(L)} \right) \|\tilde{g}(w_k, X)\|_2^2
\]

(S77)

\[
= (1 - \beta_k \lambda_2(L)) \|\tilde{w}_k\|_2^2 + \frac{2\alpha_k}{\beta_k \lambda_2(L)} \left( 2\|\tilde{v}_k\|_2^2 + n^2 \alpha_k \|\tilde{g}(w_k, X)\|_2^2 \right)
\]

(S78)
Now taking the conditional expectation $\mathbb{E} \left[ \cdot | \mathcal{F}_k \right]$ yields
\[
\mathbb{E} \left[ \| \hat{w}_{k+1} \|^2 \right] \leq (1 - \beta_k \lambda_2(\mathcal{L})) \mathbb{E} \left[ \| \hat{\tilde{w}}_k \|^2 \right] + \frac{2 \alpha_k}{\beta_k \lambda_2(\mathcal{L})} \left( 2 \mathbb{E} \left[ \| \tilde{v}_k \|^2 | \mathcal{F}_k \right] + n^2 \alpha_k \mathbb{E} \left[ \| \bar{g}(w_k, X) \|^2 | \mathcal{F}_k \right] \right)
\]
\[
\leq (1 - \beta_k \lambda_2(\mathcal{L})) \mathbb{E} \left[ \| \hat{w}_k \|^2 \right] + \frac{2 \alpha_k}{\beta_k \lambda_2(\mathcal{L})} (2n d_w + n^2 a \mu_g)
\]
(S79)

where we used Assumption 3. Taking the total expectation further gives
\[
\mathbb{E} \left[ \| \hat{w}_{k+1} \|^2 \right] \leq \left( 1 - \frac{b \lambda_2(\mathcal{L})}{1 + k \delta_1} \right) \mathbb{E} \left[ \| \hat{w}_k \|^2 \right] + \frac{2 n a (2d_w + n a \mu_g)}{b \lambda_2(\mathcal{L})} \frac{1}{(1 + k)^{\delta_2 - \delta_1}}
\]
(S80)

Now (S81) can be written in the form of (S15) with $\mu_g = b \lambda_2(\mathcal{L})$, $\delta_4 = \delta_2 - \delta_1$ and $\mu_\zeta = \frac{2 n a (2d_w + n a \mu_g)}{b \lambda_2(\mathcal{L})}$. Thus it follows from Lemma 5 that
\[
\mathbb{E} \left[ \| \hat{w}_{k+1} \|^2 \right] \leq \frac{W_3}{\exp(W_1(k + 1)^{1-\delta_1})} + \frac{W_2}{(k + 1)^{\delta_2 - \delta_1}}
\]
(S82)

where $W_1$, $W_2$ and $W_3$ are positive constants defined as
\[
W_1 = \frac{b \lambda_2(\mathcal{L})}{1 - \delta_1},
\]
(S83)
\[
W_2 = \frac{2 n a (2d_w + n a \mu_g)}{b^2 \lambda_2(\mathcal{L})^2 \delta_1} \exp(W_1 2^{1-\delta_1}),
\]
(S84)
\[
W_3 = \exp(W_1) \left( \mathbb{E} \left[ \| \hat{w}_0 \|^2 \right] + \frac{2 n a (2d_w + n a \mu_g)}{b \lambda_2(\mathcal{L})} \sum_{\ell = 0}^{k} \frac{1}{1 - b \lambda_2(\mathcal{L})} \frac{1}{(\ell + 1)^{\delta_2 - \delta_1}} \right)
\]
(S85)

in which $\lambda_2(\mathcal{L})$ denotes the second smallest eigenvalue of $\mathcal{L}$ and $\tilde{k} > 0$ is defined as
\[
\tilde{k} = \left[ \left( \frac{\delta_2 - \delta_1}{b \lambda_2(\mathcal{L})} \right)^{-1} \right].
\]
(S86)

This concludes the proof of Theorem 1.

S3.1 Consensus rate

Note that
\[
\frac{W_3}{\exp(W_1(k + 1)^{1-\delta_1})} = \frac{W_3}{(k + 1)^{\delta_2 - \delta_1}} \frac{(k + 1)^{\delta_2 - 2\delta_1}}{\exp(W_1(k + 1)^{1-\delta_1})}
\]
(S87)
\[
\leq \frac{W_3}{(k + 1)^{\delta_2 - 2\delta_1}} \max_{\ell \geq 0} \left( \frac{(t + 1)^{\delta_2 - 2\delta_1}}{\exp(W_1(t + 1)^{1-\delta_1})} \right).
\]
(S88)
We have
\[
\max_{\forall t \geq 0} \left( \frac{(t + 1)^{\delta_2 - 2\delta_1}}{\exp(W_1(t + 1)^{1-\delta_1})} \right) = \exp \left( \frac{\delta_2 - 2\delta_1}{1 - \delta_1} \right) \left( \frac{\delta_2 - 2\delta_1}{W_1(1 - \delta_1)} \right)^{\delta_2 - 2\delta_1}, \quad (S89)
\]
which is attained when
\[
\frac{\delta_2 - 2\delta_1}{1 - \delta_1} = W_1(t + 1)^{1-\delta_1}. \quad (S90)
\]
Note \( \frac{\delta_2 - 2\delta_1}{1 - \delta_1} = \frac{\delta_2 - 2\delta_1}{\lambda_2(L)} \). Define
\[
W_4 = W_3 \exp \left( \frac{\delta_2 - 2\delta_1}{1 - \delta_1} \right) \left( \frac{\delta_2 - 2\delta_1}{\lambda_2(L)} \right)^{\delta_2 - 2\delta_1}. \quad (S91)
\]
Then
\[
E \left[ \|\tilde{w}_{k+1}\|^2 \right] \leq \frac{W_2 + W_4}{(k + 1)^{\delta_2 - 2\delta_1}}. \quad (S92)
\]

### S4 Proof of Theorem 2

Denote \( \tilde{w}(t_k) \) and \( \tilde{w}(t_k) \) by \( Y_{k,1} \) and \( Y_{k,2} \), respectively. Let \( Y_k = [Y_{k,1}^T Y_{k,2}^T]^T \) and \( X_k(s) = [X_k^T \tilde{w}^T(s)]^T \). Then from (24) we have for \( s \in [t_k, t_{k+1}] \)
\[
dX_k(s) = \begin{pmatrix} 0 \\ 0 \\ -\nabla U(Y_{k,1}, X) - \zeta(Y_{k,2}, Y_{k,1}) \end{pmatrix} ds + \begin{pmatrix} 0 \\ 0 \\ \sqrt{2dB_s} \end{pmatrix}. \quad (S93)
\]
Let \( X_k \) admit a distribution \( p_t(X_k) \). The time evolution of \( p_t(X_k) \) is given by the following Fokker-Planck (FP) equation (see 4.1 in [18])

\[
\frac{\partial p_t(X_k)}{\partial t} = -\nabla \tilde{w} \cdot [p_t(X_k) (-\nabla U(Y_{k,1}, X) - \nabla \log p_t(X_k) - \zeta(Y_{k,2}, Y_{k,1}))]
\]

where \( \nabla \cdot [\tilde{u}(\cdot)] \) denotes the divergence of a vector field \( \tilde{u}(\cdot) \).

We next marginalize out \( Y_k \) from \( p_t(X_k) \) to obtain \( p_t(w) = \int p_t(X_k) dY_k \) and
\[
\frac{\partial p_t(w)}{\partial t} = -\nabla \tilde{w} \cdot \left[ \int p_t(w, Y_k) (-\nabla U(Y_{k,1}, X) - \nabla \log p_t(w, Y_k) - \zeta(Y_{k,2}, Y_{k,1})) \right] dY_k. \quad (S95)
\]
Note that
\[
\int p_t(w, Y_k) \nabla \log p_t(w, Y_k) dY_k = p_t(w) \nabla \log p_t(w) = \nabla p_t(w). \quad (S96)
\]
We further write (S95) as
\[
\frac{\partial p_t(\tilde{w})}{\partial t} = \nabla_{\tilde{w}} \cdot [p_t(\tilde{w}) \nabla \log p_t(\tilde{w})]
- \nabla_{\tilde{w}} \cdot \left[ \int p_t(\tilde{w}, Y_k) \left( -\nabla U(Y_{k,1}, X) - \zeta(Y_{k,2}, Y_{k,1}) \right) \, dY_k \right]
\]
\[
= \nabla_{\tilde{w}} \cdot [p_t(\tilde{w}) \nabla \log p_t(\tilde{w})]
- \nabla_{\tilde{w}} \cdot \left[ \int p_t(\tilde{w}, Y_k) \left( -\nabla U(\tilde{w}, X) - \zeta(Y_{k,2}, Y_{k,1}) \right) \, dY_k \right]
\]  
(S97)

\[
= \nabla_{\tilde{w}} \cdot [p_t(\tilde{w}) \nabla \log p_t(\tilde{w})] - \nabla_{\tilde{w}} \cdot \left[ \int p_t(\tilde{w}, Y_k) \left( -\nabla U(\tilde{w}, X) \right) \, dY_k \right]
\]
\[
- \nabla_{\tilde{w}} \cdot \left[ \int p_t(\tilde{w}, Y_k) (\nabla U(\tilde{w}, X) - \nabla U(Y_{k,1}, X) - \zeta(Y_{k,2}, Y_{k,1})) \, dY_k \right]
\]  
(S98)

\[
= \nabla_{\tilde{w}} \cdot [p_t(\tilde{w}) \nabla \log p_t(\tilde{w}) + \nabla U(\tilde{w}, X)]
- \nabla_{\tilde{w}} \cdot \left[ \int p_t(\tilde{w}, Y_k) (\nabla U(\tilde{w}, X) - \nabla U(Y_{k,1}, X) - \zeta(Y_{k,2}, Y_{k,1})) \, dY_k \right]
\]  
(S99)

Let
\[
f_t(\tilde{w}) = p_t(\tilde{w}) (\nabla \log p_t(\tilde{w}) + \nabla U(\tilde{w}, X)) = p_t(\tilde{w}) \nabla \log \left( \frac{p_t(\tilde{w})}{p^*(\tilde{w})} \right),
\]  
(S101)

where we used \( \nabla U(\tilde{w}, X) = -\nabla \log p^*(\tilde{w}) \) and let
\[
\tilde{f}_t(\tilde{w}) = \int p_t(\tilde{w}, Y_k) (\nabla U(\tilde{w}, X) - \nabla U(Y_{k,1}, X) - \zeta(Y_{k,2}, Y_{k,1})) \, dY_k.
\]  
(S102)

Thus,
\[
\frac{\partial p_t(\tilde{w})}{\partial t} = \nabla_{\tilde{w}} \cdot \left[ f_t - \tilde{f}_t \right].
\]  
(S103)

We next derive the evolution of the KL divergence between \( p_t(\tilde{w}) \) and \( p^*(\tilde{w}) \), denoted by \( F(p_t(\tilde{w})) \), i.e.,
\[
F(p_t(\tilde{w})) = \int p_t(\tilde{w}) \log \left( \frac{p_t(\tilde{w})}{p^*(\tilde{w})} \right) \, d\tilde{w}.
\]  
(S104)

Taking the time derivative of \( F(p_t(\tilde{w})) \) leads to
\[
\dot{F}(p_t(\tilde{w})) = \frac{d}{dt} \int p_t(\tilde{w}) \log \left( \frac{p_t(\tilde{w})}{p^*(\tilde{w})} \right) \, d\tilde{w}
\]  
(S105)
\[
= \int \frac{\partial}{\partial t} \left( p_t(\tilde{w}) \log \left( \frac{p_t(\tilde{w})}{p^*(\tilde{w})} \right) \right) \, d\tilde{w}
\]  
(S106)
The first term in (S109) corresponds to the continuous time Langevin dynamics. From Lemma S5, we further obtain

\[
\kappa(\bar{w}) = \log \left( \frac{p_t(\bar{w})}{p^*(\bar{w})} \right) + 1.
\]

Using (S103), we further obtain

\[
\tilde{F}(p_t(\bar{w})) = \int \kappa(\bar{w}) \frac{\partial p_t(\bar{w})}{\partial t} \, d\bar{w}
= \int \kappa(\bar{w}) (\nabla_{\bar{w}} \cdot [f_t(\bar{w})]) \, d\bar{w} - \int \kappa(\bar{w}) (\nabla_{\bar{w}} \cdot \left[ \tilde{f}_t(\bar{w}) \right]) \, d\bar{w}.
\]

The first term in (S109) corresponds to the continuous time Langevin dynamics. Using Lemma S5 (an alternative version of Lemma 10.4.1 in [49]) it can be shown that

\[
\int \kappa(\bar{w}) (\nabla_{\bar{w}} \cdot [f_t(\bar{w})]) \, d\bar{w} = -\int \nabla \log \left( \frac{p_t(\bar{w})}{p^*(\bar{w})} \right)^T f_t(\bar{w}) \, d\bar{w}.
\]

Substituting (S101), we further get

\[
\int \nabla \log \left( \frac{p_t(\bar{w})}{p^*(\bar{w})} \right)^T f_t(\bar{w}) \, d\bar{w} = \int \left\| \nabla \log \left( \frac{p_t(\bar{w})}{p^*(\bar{w})} \right) \right\|_2^2 p_t(\bar{w}) \, d\bar{w}
= \mathbb{E}_{p_t(\bar{w})} \left[ \left\| \nabla \log \left( \frac{p_t(\bar{w})}{p^*(\bar{w})} \right) \right\|_2^2 \right].
\]

For the second term in (S109), we have

\[
\int \kappa(\bar{w}) (\nabla_{\bar{w}} \cdot \left[ \tilde{f}_t(\bar{w}) \right]) \, d\bar{w}
= \int \kappa(\bar{w}) \left( \nabla_{\bar{w}} \cdot \left[ \int p_t(\bar{w}, Y_k) (\nabla U(\bar{w}, X) - \nabla U(Y_{k,1}, X) - \zeta(Y_{k,2}, Y_{k,1})) \, dY_k \right] \right) \, d\bar{w}
= \int \kappa(\bar{w}) (\nabla_{\bar{w}} \cdot \left[ p_t(\bar{w}, Y_k) (\nabla U(\bar{w}, X) - \nabla U(Y_{k,1}, X) - \zeta(Y_{k,2}, Y_{k,1})) \right] \, d\bar{w} \, dY_k.
\]

From Lemma S5, we further obtain

\[
\int \kappa(\bar{w}) (\nabla_{\bar{w}} \cdot \left[ p_t(\bar{w}, Y_k) (\nabla U(\bar{w}, X) - \nabla U(Y_{k,1}, X) - \zeta(Y_{k,2}, Y_{k,1})) \right] \, d\bar{w} \, dY_k
= -\int \nabla_{\bar{w}} \kappa(\bar{w})^T (\nabla U(\bar{w}, X) - \nabla U(Y_{k,1}, X) - \zeta(Y_{k,2}, Y_{k,1})) p_t(\bar{w}, Y_k) \, d\bar{w} \, dY_k
= -\int \nabla \log \left( \frac{p_t(\bar{w})}{p^*(\bar{w})} \right)^T \tilde{f}_t(\bar{w}) \, d\bar{w}.
\]
It then follows from (S109), (S112) and (S116) that
\[
\hat{F}(p_t(\bar{\omega})) = -E_{p_t(\bar{\omega})}\left[\left\|\nabla \log \left(\frac{p_t(\bar{\omega})}{p^*(\bar{\omega})}\right)\right\|_2^2 + \int \nabla \log \left(\frac{p_t(\bar{\omega})}{p^*(\bar{\omega})}\right)^\top \tilde{f}_t(\bar{\omega}) d\bar{\omega}\right].
\]
(S117)

To bound the second term in (S117), we note from (S102) that
\[
\int \nabla \log \left(\frac{p_t(\bar{\omega})}{p^*(\bar{\omega})}\right)^\top \tilde{f}_t(\bar{\omega}) d\bar{\omega} = \int \int \nabla \log \left(\frac{p_t(\bar{\omega})}{p^*(\bar{\omega})}\right)^\top (\nabla U(\bar{\omega}, X) - \nabla U(Y_{k,1}, X) - \zeta(Y_{k,2}, Y_{k,1})) p_t(\bar{\omega}, Y_k) dY_k d\bar{\omega}
\]
\[
\leq \frac{1}{2} \int \int \left\|\nabla \log \left(\frac{p_t(\bar{\omega})}{p^*(\bar{\omega})}\right)\right\|_2^2 p_t(\bar{\omega}, Y_k) dY_k d\bar{\omega}
\]
\[
+ \frac{1}{2} \int \int \left\|\nabla U(\bar{\omega}, X) - \nabla U(Y_{k,1}, X) - \zeta(Y_{k,2}, Y_{k,1})\right\|_2^2 p_t(\bar{\omega}, Y_k) dY_k d\bar{\omega}
\]
(S118)
\[
\leq \frac{1}{2} E_{p_t(\bar{\omega})}\left[\left\|\nabla \log \left(\frac{p_t(\bar{\omega})}{p^*(\bar{\omega})}\right)\right\|_2^2 + \int \|\nabla U(\bar{\omega}, X) - \nabla U(Y_{k,1}, X)\|_2^2 p_t(\bar{\omega}, Y_k) dY_k d\bar{\omega}\right]
\]
\[
+ \int \|\zeta(Y_{k,2}, Y_{k,1})\|_2^2 p_t(\bar{\omega}, Y_k) dY_k d\bar{\omega}
\]
(S119)
\[
\leq \frac{1}{2} E_{p_t(\bar{\omega})}\left[\left\|\nabla \log \left(\frac{p_t(\bar{\omega})}{p^*(\bar{\omega})}\right)\right\|_2^2 + L^2 \int \|\bar{\omega} - Y_{k,1}\|_2^2 p_t(\bar{\omega}, Y_k) dY_k d\bar{\omega}\right]
\]
\[
+ \int \|\zeta(Y_{k,2}, Y_{k,1})\|_2^2 p_t(\bar{\omega}, Y_k) dY_k d\bar{\omega}
\]
(S120)

For the last term in (S121), it follows from Assumption [1] that
\[
\left\|\zeta(\bar{\omega}(t_k), \bar{\omega}(t_k))\right\|_2 \leq \sum_{i=1}^n \left\|\nabla U_i(\bar{\omega}(t_k)) + \zeta_i(t_k), X_i) - \nabla U_i(\bar{\omega}(t_k), X_i)\right\|_2 \leq L^2 \left\|\bar{\omega}(t_k)\right\|_2^2
\]
(S121)

from which we obtain
\[
\int \int \|\zeta(Y_{k,2}, Y_{k,1})\|_2^2 p_t(\bar{\omega}, Y_k) dY_k d\bar{\omega} \leq L^2 \int \|Y_{k,2}\|_2^2 p_t(\bar{\omega}, Y_k) dY_k d\bar{\omega}
\]
(S122)
\[
= L^2 \int \|Y_{k,2}\|_2^2 p_t(\bar{\omega}, Y_{k,2}) dY_{k,2} d\bar{\omega}
\]
(S123)

\[
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\]
\[ L^2 \int \|Y_{k,2}\|^2 p(Y_{k,2}) \, dY_{k,2} \] (S125)

\[ = L^2 \mathbb{E}_{p(Y_{k,2})} \|Y_{k,2}\|^2 = L^2 \mathbb{E}_{p(t)} \left( \|\hat{w}(t)\|^2 \right). \] (S126)

For the second term in \((S121)\), since \(Y_{k,1} = \tilde{w}(t_k)\), it follows from \((S93)\) that \(\forall t \in [t_k, t_{k+1}]\)

\[ \|\hat{w}(t) - \hat{w}(t_k)\|^2 \geq \left\| \nabla U(\hat{w}(t), x)(t - t_k) + \sqrt{2}(B_t - B_{t_k}) - (t - t_k)\zeta(\hat{w}(t), \hat{w}(t_k)) \right\|^2 \] (S127)

\[ = 2\|B_t - B_{t_k}\|^2 + \|\nabla U(\hat{w}(t), x)(t - t_k) + (t - t_k)\zeta(\hat{w}(t), \hat{w}(t_k))\|^2 \]

\[ - 2\sqrt{2}(B_t - B_{t_k})^\top (\nabla U(\hat{w}(t), x)(t - t_k) + (t - t_k)\zeta(\hat{w}(t), \hat{w}(t_k))) \] (S128)

\[ \leq 2\|B_t - B_{t_k}\|^2 + 2\sigma_e^2 \|\nabla U(\hat{w}(t), x)(t - t_k) + (t - t_k)\zeta(\hat{w}(t), \hat{w}(t_k))\|^2 \]

\[ - 2\sqrt{2}(B_t - B_{t_k})^\top (\nabla U(\hat{w}(t), x)(t - t_k) + (t - t_k)\zeta(\hat{w}(t), \hat{w}(t_k))) \] (S129)

\[ \leq 2\|B_t - B_{t_k}\|^2 + 2\sigma_e^2 L^2 \|\hat{w}(t_k)\|^2 + 2\sigma_e^2 L^2 \|\hat{w}(t)\|^2 \]

\[ - 2\sqrt{2}(B_t - B_{t_k})^\top (\nabla U(\hat{w}(t), x)(t - t_k) + (t - t_k)\zeta(\hat{w}(t), \hat{w}(t_k))). \] (S130)

The last inequality follows from Lipschitz continuity of \(\nabla U\) and assuming \(\nabla U(0, x) = 0\). Note that assuming \(\nabla U(0, x) = 0\) is only to simplify the notation. Later we will bound the expectation of \(\|\hat{w}(t)\|^2\) in \((S130)\). Then given any finite \(w^*\) such that \(\nabla U(w^*, x) = 0\), \(\|\nabla U(\hat{w}(t), x)\|^2 \leq L^2 \|\hat{w}(t) - w^*\|^2\), whose expectation is also bounded.

Let \(\tilde{B}_t = B_t - B_{t_k}\). Then \(\tilde{B}_t\) follows a zero mean Gaussian distribution with a variance of \(t - t_k\). Note that \(\hat{w}(t)\) depends on \(\tilde{B}_t\) for \(t_k < t \leq t_{k+1}\) while \(Y_k\) is independent of \(\tilde{B}_t\). Thus,

\[ \int \int \|\tilde{B}_t\|^2 p_t(\hat{w}, Y_k) \, dY_k \, d\hat{w} = \int \|\tilde{B}_t\|^2 p_t(\hat{w}) \, d\hat{w} \] (S131)

\[ = \int \|\tilde{B}_t\|^2 \left( \int p_t(\hat{w}) \tilde{B}_t \, p(\tilde{B}_t) \, d\tilde{B}_t \right) \, d\hat{w} \] (S132)

\[ = \int \|\tilde{B}_t\|^2 p(\tilde{B}_t) \left( \int p_t(\hat{w}) \tilde{B}_t \, d\tilde{B}_t \right) \, d\tilde{B}_t \] (S133)

\[ = E_{p(\tilde{B}_t)} \|\tilde{B}_t\|^2 = d_w(t - t_k) \leq \alpha_k d_w. \] (S134)

Similarly, for any function \(S\) of \(Y_k\), we have

\[ \int \int \tilde{B}_t^\top S(Y_k) p_t(\hat{w}, Y_k) \, dY_k \, d\hat{w} = \int \int \tilde{B}_t^\top S(Y_k) \left( \int p_t(\hat{w} | Y_k, \tilde{B}_t) \, p(Y_k) \, p(\tilde{B}_t) \, d\tilde{B}_t \right) \, dY_k \, d\hat{w} \] (S135)

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Now substituting (S141) into (S117) gives
\[
\int \sigma_Y(Y_k)p(Y_k)p(B_t) \left[ \int p_t(\bar{w}|Y_k, \bar{B}_t) \; d\bar{w} \right] \; dY_k \; d\bar{B}_t
\]
which is
\[
= \int p(B_t)\bar{B}_t^T S(Y_k)p(Y_k) \; dY_k \; d\bar{B}_t = 0. \quad \text{(S137)}
\]
Recall \( Y_{k,1} = \bar{w}(t_k) \) and \( Y_{k,2} = \tilde{\bar{w}}(t_k) \). Using (S134) and (S137), we further obtain
\[
\int \int \| \bar{w}(t) - Y_{k,1} \|^2 \; p_t(\bar{w}, Y_k) \; dY_k \; d\bar{w}
\]
\[
\leq \int \int \left( 2\alpha_k^2 L^2 \| Y_{k,1} \|^2 + 2\alpha_k^2 L^2 \| Y_{k,2} \|^2 + 2\alpha_k d_w \right) p_t(\bar{w}, Y_k) \; dY_k \; d\bar{w}
\]
\[
= 2\alpha_k^2 L^2 E_{p(Y_{k,1})} \| Y_{k,1} \|^2 + 2\alpha_k^2 L^2 E_{p(Y_{k,2})} \| Y_{k,2} \|^2 + 2\alpha_k d_w
\]
\[
= 2\alpha_k^2 L^2 E_{p(\bar{w}(t_k))} \| \bar{w}(t_k) \|^2 + 2\alpha_k^2 L^2 E_{p(\tilde{\bar{w}}(t_k))} \| \tilde{\bar{w}}(t_k) \|^2 + 2\alpha_k d_w. \quad \text{(S139)}
\]
Substituting (S140) and (S126) into (S121) yields
\[
\int \nabla \log \left( \frac{p_t(\bar{w})}{p_t(\bar{w})} \right)^T \tilde{f}_t(\bar{w}) \; d\bar{w} \leq \frac{1}{2} E_{p_t(\bar{w})} \left[ \left\| \nabla \log \left( \frac{p_t(\bar{w})}{p_t(\bar{w})} \right) \right\|^2_t + L^2 E_{p(\tilde{\bar{w}}(t_k))} \| \tilde{\bar{w}}(t_k) \|_2^2 \right]
\]
\[
+ \tilde{L}^2 \left( 2\alpha_k^2 L^2 E_{p(\bar{w}(t_k))} \| \bar{w}(t_k) \|^2 + 2\alpha_k^2 L^2 E_{p(\tilde{\bar{w}}(t_k))} \| \tilde{\bar{w}}(t_k) \|^2 + 2\alpha_k d_w \right). \quad \text{(S141)}
\]
Now substituting (S141) into (S117) gives
\[
\tilde{F}(p_t(\bar{w})) \leq -\frac{1}{2} E_{p_t(\bar{w})} \left[ \left\| \nabla \log \left( \frac{p_t(\bar{w})}{p_t(\bar{w})} \right) \right\|^2_t + L^2 E_{p(\tilde{\bar{w}}(t_k))} \| \tilde{\bar{w}}(t_k) \|_2^2 \right]
\]
\[
+ \tilde{L}^2 \left( 2\alpha_k^2 L^2 E_{p(\bar{w}(t_k))} \| \bar{w}(t_k) \|^2 + 2\alpha_k^2 L^2 E_{p(\tilde{\bar{w}}(t_k))} \| \tilde{\bar{w}}(t_k) \|^2 + 2\alpha_k d_w \right) \quad \text{(S142)}
\]
\[
= -\frac{1}{2} E_{p_t(\bar{w})} \left[ \left\| \nabla \log \left( \frac{p_t(\bar{w})}{p_t(\bar{w})} \right) \right\|^2_t + \left( 2\alpha_k^2 L^2 L^2 + L^2 \right) E_{p(\tilde{\bar{w}}(t_k))} \| \tilde{\bar{w}}(t_k) \|_2^2 \right]
\]
\[
+ 2\alpha_k^2 L^4 E_{p(\bar{w}(t_k))} \| \bar{w}(t_k) \|^2 + 2\alpha_k L \tilde{L} d_w. \quad \text{(S143)}
\]
Recall the log-Sobolev inequality (21)
\[
F(p_t(\bar{w})) = E_{p_t(\bar{w})} \left[ \log \left( \frac{p_t(\bar{w})}{p_t(\bar{w})} \right) \right] \leq \frac{1}{2\rho_U} E_{p_t(\bar{w})} \left[ \left\| \nabla \log \left( \frac{p_t(\bar{w})}{p_t(\bar{w})} \right) \right\|^2_t \right],
\]
\[
\text{(S144)}
\]
where $\rho_U$ is the log-Sobolev constant. We then have
\[
\dot{F}(p_t(\bar{w})) \leq -\rho_U F(p_t(\bar{w})) + 2\alpha_k^2 L^4 \mathbb{E}_{p(\bar{w}(t_k))} \|\bar{w}(t_k)\|_2^2 \\
+ (2\alpha_k^2 L^2 L^2 + L^2) \mathbb{E}_{p(\bar{w}(t_k))} \|\hat{w}(t_k)\|_2^2 + 2\alpha_k L^2 d_w
\]
(S145)

which means $\forall t \in [t_k, t_{k+1}]
\[
F(p_t(\bar{w})) \leq \exp(-\rho_U(t - t_k)) F(p_{t_k}(\bar{w})) + \frac{1 - \exp(-\rho_U(t - t_k))}{\rho_U} \left( 2\alpha_k^2 L^4 \mathbb{E}_{p(\bar{w}(t_k))} \|\bar{w}(t_k)\|_2^2 \\
+ (2\alpha_k^2 L^2 L^2 + L^2) \mathbb{E}_{p(\bar{w}(t_k))} \|\hat{w}(t_k)\|_2^2 + 2\alpha_k L^2 d_w \right).
\]
(S146)

Since $\frac{1 - \exp(-\rho_U(t - t_k))}{\rho_U} \leq t - t_k \leq \alpha_k$, we further obtain
\[
F(p_t(\bar{w})) \leq \exp(-\rho_U(t - t_k)) F(p_{t_k}(\bar{w})) + \left( 2\alpha_k^2 L^4 \mathbb{E}_{p(\bar{w}(t_k))} \|\bar{w}(t_k)\|_2^2 \\
+ (2\alpha_k^2 L^2 L^2 + L^2) \mathbb{E}_{p(\bar{w}(t_k))} \|\hat{w}(t_k)\|_2^2 + 2\alpha_k L^2 d_w \right).
\]
(S147)

In particular, at $t = t_{k+1}$, we have
\[
F(p_{t_{k+1}}(\bar{w})) \leq \exp(-\rho_U\alpha_k) F(p_{t_k}(\bar{w})) + \left( 2\alpha_k^2 L^4 \mathbb{E}_{p(\bar{w}(t_k))} \|\bar{w}(t_k)\|_2^2 \\
+ (2\alpha_k^2 L^2 L^2 + L^2) \mathbb{E}_{p(\bar{w}(t_k))} \|\hat{w}(t_k)\|_2^2 + 2\alpha_k L^2 d_w \right).
\]
(S149)

We can then use (S149) to recursively bound the KL-divergence $F(p_{t_{k+1}}(\bar{w}))$.

Note that $\mathbb{E}_{p(\bar{w}(t_k))} \|\bar{w}(t_k)\|_2^2 \leq C_w$ from Lemma [S9] if we select $a$ in $\alpha_k = \frac{a}{(k+1)^2}$ as
\[
a = \frac{1}{n^{4/3}} \left( \frac{\rho_U (3\delta_2 - 1)}{25 L^4 \delta_2} \right)^{4/3}.
\]
(S150)

Also recall from Theorem [1] that
\[
\mathbb{E} \left[ \|\bar{w}_{k+1}\|_2^2 \right] \leq \frac{W_3}{\exp(W_1(k + 1)^{1-\delta_1})} + \frac{W_2}{(k+1)^{\delta_2 - 2\delta_1}}.
\]
(S151)
Let
\[ Z_k = \frac{W_3(2\alpha_k^3\bar{L}^2L^2 + L^2\alpha_k)}{\exp(W_1(k+1)^{1-\delta_1})}, \]  
(S152)
\[ C_\alpha = \frac{\rho_U(3\delta_2 - 1)}{25L^4\delta_2}, \]  
(S153)
and
\[ \xi_k = 2\alpha_k^3\bar{L}^4\bar{C}_w + (2\alpha_k^3\bar{L}^2L^2 + L^2\alpha_k)\frac{W_2}{(k+1)^{2\delta_2 - 2\delta_1}} + 2\alpha_k^3\bar{L}^2\bar{d}_w + Z_k. \]  
(S154)

Also define
\[ \theta_k = \frac{\bar{W}_3}{\exp(W_1(k+1)^{1-\delta_1})}, \]  
(S155)
where
\[ \bar{W}_3 = W_3(2\alpha^3\bar{L}^2L^2 + L^2a). \]  
(S156)

Note that \( Z_k \leq \theta_k \). Now substituting
\[ \alpha_k = \left(\frac{C_\alpha}{n^4}\right)^\frac{1}{2} \frac{1}{(k+1)^{\delta_2}} \]  
(S157)
into (S153) yields
\[ \xi_k = \frac{2C_\alpha \bar{L}^4\bar{C}_w}{n^4(k+1)^{3\delta_2}} + \left(\frac{2C_\alpha \bar{L}^2L^2}{n^4(k+1)^{3\delta_2}} + \frac{C_\alpha^\frac{1}{2} \bar{L}^2}{n^{\frac{1}{2}}(k+1)^{\frac{1}{2}\delta_2}}\right)\frac{W_2}{(k+1)^{2\delta_2 - 2\delta_1}} + \frac{2C_\alpha^\frac{3}{4} \bar{L}^2\bar{d}_w}{n^{\frac{1}{4}}(k+1)^{2\delta_2}} + Z_k \]  
(S158)
\[ = \frac{2C_\alpha \bar{L}^4\bar{C}_w}{n^4(k+1)^{3\delta_2}} + \frac{2C_\alpha \bar{L}^2L^2W_2}{n^4(k+1)^{4\delta_2 - 2\delta_1}} + \frac{C_\alpha^\frac{3}{2} \bar{L}^2W_2}{n^{\frac{3}{4}}(k+1)^{2\delta_2 - 2\delta_1}} + \frac{2C_\alpha^\frac{3}{4} \bar{L}^2\bar{d}_w}{n^{\frac{1}{4}}(k+1)^{2\delta_2}} + Z_k \]  
(S159)
\[ \leq \frac{C_\xi}{(k+1)^{2\delta_2 - 2\delta_1}} + \theta_k, \]  
(S160)
where
\[ C_\xi = \frac{C_\alpha^\frac{1}{2}}{n^{\frac{1}{4}}} \left(\frac{2C_\alpha \bar{L}^4\bar{C}_w}{n^{\frac{1}{4}}} + \frac{2C_\alpha \bar{L}^2L^2W_2}{n^{\frac{3}{4}}} + \frac{C_\alpha^\frac{3}{2} \bar{L}^2W_2}{n^{\frac{5}{4}}(k+1)^{2\delta_2}} + \frac{2C_\alpha^\frac{3}{4} \bar{L}^2\bar{d}_w}{n^{\frac{1}{4}}(k+1)^{2\delta_2}}\right). \]  
(S161)

Now we rewrite (S149) as
\[ F(p_{t_k+1}(\bar{w})) \leq \exp(-\rho_U \alpha_k) F(p_{t_k}(\bar{w})) + \xi_k, \]  
(S162)
which results in
\[
F(p_{t_{k+1}}(\bar{w})) \leq F(p_{t_0}(\bar{w})) \exp \left( -\rho_U \sum_{\ell=0}^{k} \alpha_\ell \right) + \sum_{\ell=0}^{k} \xi_\ell \exp \left( -\rho_U \sum_{i=\ell+1}^{k} \alpha_i \right).
\]
(S163)

When \( \delta_2 \in (0.5, 1) \), from Lemma S3 we have
\[
\sum_{\ell=0}^{k} \alpha_\ell \geq \int_{0}^{k} \frac{a}{(x+1)^{\delta_2}} \, dx = \frac{a(k+1)^{1-\delta_2}}{1-\delta_2} - \frac{a}{1-\delta_2}.
\]
(S164)

Thus
\[
\exp \left( -\rho_U \sum_{\ell=0}^{k} \alpha_\ell \right) \leq \exp \left( -\frac{\rho_U a}{1-\delta_2} (k+1)^{1-\delta_2} + \frac{\rho_U a}{1-\delta_2} \right)
\]
(S165)
\[
= \exp \left( -\frac{\rho_U a}{1-\delta_2} (k+1)^{1-\delta_2} \right) \exp \left( \frac{\rho_U a}{1-\delta_2} \right).
\]
(S166)

Therefore from (S163) we obtain
\[
F(p_{t_{k+1}}(\bar{w})) \leq F(p_{t_0}(\bar{w})) \exp \left( \frac{\rho_U a}{1-\delta_2} \right) \exp \left( -\frac{\rho_U a}{1-\delta_2} (k+1)^{1-\delta_2} \right) + \sum_{\ell=0}^{k} \xi_\ell \exp \left( -\rho_U \sum_{i=\ell+1}^{k} \alpha_i \right).
\]
(S167)

Substituting (S160) yields
\[
F(p_{t_{k+1}}(\bar{w})) \leq F(p_{t_0}(\bar{w})) \exp \left( \frac{\rho_U a}{1-\delta_2} \right) \exp \left( -\frac{\rho_U a}{1-\delta_2} (k+1)^{1-\delta_2} \right) + \sum_{\ell=0}^{k} \frac{C_\xi}{(\ell+1)^{2\delta_2-2\delta_1}} \exp \left( -\rho_U \sum_{i=\ell+1}^{k} \alpha_i \right) + \sum_{\ell=0}^{k} \theta_\ell \exp \left( -\rho_U \sum_{i=\ell+1}^{k} \alpha_i \right).
\]
(S168)

We first consider the last term in (S168). We write
\[
\theta_\ell \exp \left( -\rho_U \sum_{i=\ell+1}^{k} \alpha_i \right) = \theta_\ell \exp(\rho_U a) \exp \left( -\rho_U \sum_{i=\ell}^{k} \alpha_i \right) \leq \theta_\ell \exp(\rho_U a) \exp \left( -\rho_U \sum_{i=\ell}^{k} \alpha_i \right).
\]
(S169)

From Lemma S3 we have
\[
\sum_{i=\ell}^{k} \alpha_i \geq \int_{\ell}^{k} \frac{a}{(x+1)^{\delta_2}} \, dx = \frac{a(k+1)^{1-\delta_2}}{1-\delta_2} - \frac{a(\ell+1)^{1-\delta_2}}{1-\delta_2} = \frac{a ((k+1)^{1-\delta_2} - (\ell+1)^{1-\delta_2})}{1-\delta_2}.
\]
(S170)
Using (S155), (S169) and (S170), we bound the last term in (S168) by

\[ \sum_{\ell=0}^{k} \theta_{\ell} \exp \left( -\rho U \sum_{i=\ell+1}^{k} \alpha_{i} \right) \leq \sum_{\ell=0}^{k} \bar{W}_3 \exp(\rho U a) \exp \left( -\frac{\rho U a ((k+1)^{1-\delta_2} - k^1)}{1 - \delta_2} \right) \]

(S171)

\[ = \bar{W}_3 \exp(\rho U a) \exp \left( -\frac{\rho U a}{1 - \delta_2} (k + 1)^{1-\delta_2} \right) \sum_{\ell=0}^{k} \exp \left( \frac{\rho U a}{1 - \delta_2} (\ell + 1)^{1-\delta_2} - W_1 (\ell + 1)^{1-\delta_1} \right) \]

(S172)

\[ = \bar{W}_3 \exp(\rho U a) \exp \left( -\frac{\rho U a}{1 - \delta_2} (k + 1)^{1-\delta_2} \right) \times \sum_{\ell=0}^{k} \exp \left( (\ell + 1)^{1-\delta_1} \left( \frac{\rho U a}{1 - \delta_2} (\ell + 1)^{1-\delta_2} - W_1 \right) \right) . \]

(S173)

Since \( \delta_1 < \delta_2 \), \( \frac{\rho U a}{1 - \delta_2} (\ell + 1)^{1-\delta_2} - W_1 \) is decreasing. There exists a finite integer \( \bar{\ell} \) such that \( \frac{\rho U a}{1 - \delta_2} (\ell + 1)^{1-\delta_2} - W_1 \geq 0 \), \( \forall \ell \leq \bar{\ell} \), and \( \frac{\rho U a}{1 - \delta_2} (\ell + 1)^{1-\delta_2} - W_1 < 0 \), \( \forall \ell > \bar{\ell} \). The sequence \( \exp \left( (\ell + 1)^{1-\delta_1} \left( \frac{\rho U a}{1 - \delta_2} (\ell + 1)^{1-\delta_2} - W_1 \right) \right) \) is decreasing after \( \bar{\ell} \) because

\[ \frac{d}{d\ell} \left[ (\ell + 1)^{1-\delta_1} \left( \frac{\rho U a}{1 - \delta_2} (\ell + 1)^{1-\delta_2} - W_1 \right) \right] < 0, \quad \forall \ell > \bar{\ell}. \]

(S174)

Note that \( \bar{\ell} \) can be computed as

\[ \bar{\ell} = \left( \frac{\rho U a}{(1 - \delta_2)W_1} \right)^{\frac{1}{2 - \delta_1}} - 1. \]

(S175)

It then follows that for \( k > \bar{\ell} \)

\[ \sum_{\ell=0}^{k} \exp \left( \frac{\rho U a}{1 - \delta_2} (\ell + 1)^{1-\delta_2} - W_1 (\ell + 1)^{1-\delta_1} \right) = \]

\[ \sum_{\ell=0}^{\bar{\ell}} \exp \left( \frac{\rho U a}{1 - \delta_2} (\ell + 1)^{1-\delta_2} - W_1 (\ell + 1)^{1-\delta_1} \right) + \sum_{\ell=\bar{\ell}+1}^{k} \exp \left( \frac{\rho U a}{1 - \delta_2} (\ell + 1)^{1-\delta_2} - W_1 (\ell + 1)^{1-\delta_1} \right) . \]

(S176)
Applying Lemma S3 yields
\[ \sum_{\ell=\ell+1}^k \exp \left( \frac{\rho U a}{1 - \delta_2} (\ell + 1)^{1-\delta_2} - W_1 (\ell + 1)^{1-\delta_1} \right) \leq \] (S177)
\[
\int_{\ell}^k \exp \left( (t + 1)^{1-\delta_1} \left( \frac{\rho U a}{1 - \delta_2} (t + 1)^{1-\delta_2} - W_1 \right) \right) dt. \] (S178)

Let \( \kappa = - \left( \frac{\rho U a}{1 - \delta_2} (\ell + 1)^{1-\delta_2} - W_1 \right) \). Note \( \kappa > 0 \). Then
\[
\int_{\ell}^k \exp \left( (t + 1)^{1-\delta_1} \left( \frac{\rho U a}{1 - \delta_2} (t + 1)^{1-\delta_2} - W_1 \right) \right) dt \leq \int_{\ell}^k \exp \left( -\kappa (t + 1)^{1-\delta_1} \right) dt 
\] (S179)
\[
= \kappa^{-1/1} \frac{1}{1 - \delta_1} \int_{\kappa(\ell+1)^{1-\delta_1}}^{\kappa(k+1)^{1-\delta_1}} \exp (-z) z^{1/\delta_1} dz \] (S180)
\[
\leq \kappa^{-1/1} \frac{1}{1 - \delta_1} \int_{0}^{\infty} \exp (-z) z^{1/\delta_1} dz \] (S181)
\[
\leq \kappa^{-1/1} \frac{1}{1 - \delta_1} \Gamma \left( \frac{1}{1 - \delta_1} \right) \] (S182)

where \( \Gamma(\cdot) \) is the Gamma function defined as
\[ \Gamma(z) = \int_{0}^{\infty} x^{z-1} \exp(-x) dx, \quad \forall z > 0. \] (S183)

We substitute (S170) into (S173) together with the bound in (S182) to get
\[
\sum_{\ell=0}^k \theta_{\ell} \exp \left( -\rho U \sum_{i=\ell+1}^k \alpha_i \right) \leq C_\theta \exp \left( -\frac{\rho U a}{1 - \delta_2} (k + 1)^{1-\delta_2} \right) \] (S184)

where
\[
C_\theta = W_3 \exp(\rho U a) \left( \sum_{\ell=0}^{\hat{\ell}} \exp \left( \frac{\rho U a}{1 - \delta_2} (\ell + 1)^{1-\delta_2} - W_1 (\ell + 1)^{1-\delta_1} \right) \right) \] (S185)
\[
+ \kappa^{-1/1} \frac{1}{1 - \delta_1} \Gamma \left( \frac{1}{1 - \delta_1} \right). \]

We next consider the second term on the right hand side of (S183). Note
that for some $\tilde{k} \in (0, k)$, we have

\[
\sum_{\ell=0}^{k} \frac{C_\xi}{(\ell + 1)^{2\delta_2 - 2\delta_1}} \exp \left( -\rho U \sum_{i=\ell+1}^{k} \alpha_i \right) = \sum_{\ell=0}^{\tilde{k}} \frac{C_\xi}{(\ell + 1)^{2\delta_2 - 2\delta_1}} \exp \left( -\rho U \sum_{i=\ell+1}^{k} \alpha_i \right) + \sum_{\ell=\tilde{k}+1}^{k} \frac{C_\xi}{(\ell + 1)^{2\delta_2 - 2\delta_1}} \exp \left( -\rho U \sum_{i=\ell+1}^{k} \alpha_i \right).
\]

From Lemma S3, we have

\[
\sum_{i=\ell+1}^{k} \alpha_i \geq \int_{\ell+1}^{k} \frac{a}{(x + 1)^{\delta_2}} dx = \frac{a(k + 1)^{1 - \delta_2}}{1 - \delta_2} - \frac{a(\ell + 2)^{1 - \delta_2}}{1 - \delta_2} = \frac{a \left( (k + 1)^{1 - \delta_2} - (\ell + 2)^{1 - \delta_2} \right)}{1 - \delta_2}.
\]

The second term in (S186) is bounded as follows

\[
\sum_{\ell=k+1}^{k} \frac{C_\xi}{(\ell + 1)^{2\delta_2 - 2\delta_1}} \exp \left( -\rho U \sum_{i=\ell+1}^{k} \alpha_i \right) \leq \sum_{\ell=k+1}^{k} \frac{C_\xi \exp \left( -\rho U \frac{a(k + 1)^{1 - \delta_2}}{1 - \delta_2} \right)}{(\ell + 1)^{2\delta_2 - 2\delta_1} \frac{a(\ell + 2)^{1 - \delta_2}}{1 - \delta_2}}
\]

(S188)

\[
= \sum_{\ell=k+1}^{k} \frac{C_\xi \exp \left( -\rho U \frac{a(k + 1)^{1 - \delta_2}}{1 - \delta_2} + \frac{\rho U a}{1 - \delta_2}(\ell + 2)^{1 - \delta_2} \right)}{(\ell + 1)^{2\delta_2 - 2\delta_1} \frac{a(\ell + 2)^{1 - \delta_2}}{1 - \delta_2}}
\]

(S189)

\[
\leq \sum_{\ell=k+1}^{k} \frac{C_\xi \exp \left( -\rho U \frac{a(k + 1)^{1 - \delta_2}}{1 - \delta_2} + \frac{\rho U a}{1 - \delta_2}2^{1 - \delta_2} \right)}{(\ell + 1)^{2\delta_2 - 2\delta_1} \frac{a(\ell + 2)^{1 - \delta_2}}{1 - \delta_2}}
\]

(S190)

\[
\leq \sum_{\ell=k+1}^{k} \frac{C_\xi \exp \left( -\rho U \frac{a(k + 1)^{1 - \delta_2}}{1 - \delta_2} + \frac{\rho U a}{1 - \delta_2}2^{1 - \delta_2} \right)}{\ell^{2\delta_2 - 2\delta_1}}
\]

(S191)

Note that from Lemma S3, we have

\[
\sum_{\ell=k+1}^{k} \frac{\exp \left( -\frac{\rho U a}{1 - \delta_2} \ell^{1 - \delta_2} \right)}{\ell^{2\delta_2 - 2\delta_1}} \leq \int_{k}^{k+1} \frac{\exp \left( -\frac{\rho U a}{1 - \delta_2} \ell^{1 - \delta_2} \right)}{\ell^{2\delta_2 - 2\delta_1}} d\ell.
\]

(S193)
Now it follows from the proof of Lemma [S4](see [S47]) that for \( \tilde{k} = \left( \frac{2\delta_2 - \delta_1}{\rho U a} \right)^\frac{1}{2\delta_2 - 2\delta_1} \),
we have
\[
\int_{k}^{k+1} \frac{\exp \left( \frac{\rho_U a}{1 - \delta_2} (1 - \delta_2) \right)}{t^{2\delta_2 - 2\delta_1}} \, dt \leq 2\delta_2 - 2\delta_1 \left( \frac{\exp \left( \frac{\rho_U a}{1 - \delta_2} (1 - \delta_2) \right)}{t^{2\delta_2 - 2\delta_1}} \right)^{k+1} \left( \frac{\exp \left( \frac{\rho_U a}{1 - \delta_2} (k + 1) (1 - \delta_2) \right)}{k^{2\delta_2 - 2\delta_1}} \right) \left( \frac{\exp \left( \frac{\rho_U a}{1 - \delta_2} (k + 1) (1 - \delta_2) \right)}{k^{2\delta_2 - 2\delta_1}} \right)
\]
\[
= \frac{2\delta_2 - 2\delta_1}{\rho U a \delta_2} \left( \frac{\exp \left( \frac{\rho_U a}{1 - \delta_2} (k + 1) (1 - \delta_2) \right)}{(k + 1)^{2\delta_2 - 2\delta_1}} - \frac{\exp \left( \frac{\rho_U a}{1 - \delta_2} (k + 1) (1 - \delta_2) \right)}{k^{2\delta_2 - 2\delta_1}} \right).
\]
We further obtain
\[
C_\xi \exp \left( -\frac{\rho_U a}{1 - \delta_2} (k + 1) (1 - \delta_2) + \frac{\rho_U a}{1 - \delta_2} 2(1 - \delta_2) \right) \sum_{\ell=k+1}^{k} \frac{\exp \left( \frac{\rho_U a}{1 - \delta_2} (1 - \delta_2) \right)}{\ell^{2\delta_2 - 2\delta_1}} \leq C_\xi \exp \left( -\frac{\rho_U a}{1 - \delta_2} (k + 1) (1 - \delta_2) + \frac{\rho_U a}{1 - \delta_2} 2(1 - \delta_2) \right)
\times \frac{2\delta_2 - 2\delta_1}{\rho U a \delta_2} \left( \frac{\exp \left( \frac{\rho_U a}{1 - \delta_2} (k + 1) (1 - \delta_2) \right)}{(k + 1)^{2\delta_2 - 2\delta_1}} - \frac{\exp \left( \frac{\rho_U a}{1 - \delta_2} (k + 1) (1 - \delta_2) \right)}{k^{2\delta_2 - 2\delta_1}} \right)
\]
\[
= \frac{2C_\xi (\delta_2 - \delta_1)}{\rho U a \delta_2} \frac{\exp \left( \frac{\rho_U a}{1 - \delta_2} 2(1 - \delta_2) \right)}{(k + 1)^{2\delta_2 - 2\delta_1}} \left( \frac{\exp \left( \frac{\rho_U a}{1 - \delta_2} (k + 1) (1 - \delta_2) \right)}{(k + 1)^{2\delta_2 - 2\delta_1}} - \frac{\exp \left( \frac{\rho_U a}{1 - \delta_2} (k + 1) (1 - \delta_2) \right)}{k^{2\delta_2 - 2\delta_1}} \right)
\]
\[
= \frac{2C_\xi (\delta_2 - \delta_1)}{\rho U a \delta_2} \frac{\exp \left( \frac{\rho_U a}{1 - \delta_2} 2(1 - \delta_2) \right)}{(k + 1)^{2\delta_2 - 2\delta_1}} \left( \frac{\exp \left( \frac{\rho_U a}{1 - \delta_2} (k + 1) (1 - \delta_2) \right)}{k^{2\delta_2 - 2\delta_1}} \right).
\]
Therefore, using [S183], [S189], and [S199], we rewrite [S168] as
\[
F(p_{t_k+1}(\overline{w})) \leq F(p_{t_0}(\overline{w})) \exp \left( -\rho_U \sum_{\ell=0}^{k} \alpha_\ell \right) + C_\theta \exp \left( -\frac{\rho_U a}{1 - \delta_2} (k + 1) (1 - \delta_2) \right)
\]
\[
+ \sum_{\ell=0}^{k} \frac{C_\xi}{(\ell + 1)^{2\delta_2 - 2\delta_1}} \exp \left( -\rho_U \sum_{i=\ell+1}^{k} \alpha_i \right) + \sum_{\ell=k+1}^{k} \frac{C_\xi}{(\ell + 1)^{2\delta_2 - 2\delta_1}} \exp \left( -\rho_U \sum_{i=\ell+1}^{k} \alpha_i \right)
\]
\[
(S200)
\]
where the last inequality follows from

Note that

\[
\sum_{i=0}^{k} \exp \left( -\rho U \sum_{i=0}^{\ell} \alpha_i \right) \leq \sum_{i=0}^{k} \frac{C_\xi \exp (\rho U \sum_{i=0}^{\ell} \alpha_i)}{\ell + 1)^{2\delta_2 - 2\delta_1}} \leq \sum_{i=0}^{k} C_\xi \exp (\rho U a \exp \left( \rho U \sum_{i=1}^{\ell} \alpha_i \right)) \frac{\exp \left( \frac{\rho U a}{1 - \delta_2} (\ell^{1-\delta_2}) \right)}{(\ell + 1)^{2\delta_1}}. \tag{S206}
\]

where the last inequality follows from

\[
\sum_{i=1}^{\ell} \alpha_i \leq \int_{0}^{\ell} \frac{a}{(x + 1)^{\delta_2}} \, dx = \frac{a}{1 - \delta_2} \left( (\ell + 1)^{1-\delta_2} - 1 \right) \leq \frac{a}{1 - \delta_2} (\ell^{1-\delta_2} + 1^{1-\delta_2} - 1) = \frac{a}{1 - \delta_2} (\ell^{1-\delta_2}). \tag{S208}
\]
Therefore (S203) can be written as

\[
F(p_{k+1}(\bar{w})) \leq \left( F(p_0(\bar{w})) + \sum_{\ell=0}^{k} C_\xi \exp \left( \frac{\rho U}{1-\delta_2} \left( \ell^{1-\delta_2} \right) \right) \exp \left( -\rho U \sum_{\ell=0}^{k} \alpha_\ell \right) \right) + C_\theta \exp \left( -\frac{\rho U a}{1-\delta_2} (k+1)^{1-\delta_2} \right) + \frac{2 C_\zeta (\delta_2 - \delta_1) \exp \left( \frac{\rho U a}{1-\delta_2} 2^{1-\delta_2} \right)}{(k+1)^{\delta_2 - 2\delta_1}}. \tag{S209}
\]

Define \( C_\rho \) as

\[
C_\rho = \frac{2 (\delta_2 - \delta_1)}{\rho U \delta_2} \exp \left( \frac{\rho U a}{1-\delta_2} 2^{1-\delta_2} \right). \tag{S210}
\]

Now (S209) can be written as

\[
F(p_{k+1}(\bar{w})) \leq \left( F(p_0(\bar{w})) + \sum_{\ell=0}^{k} C_\xi \exp \left( \frac{\rho U}{1-\delta_2} \left( \ell^{1-\delta_2} \right) \right) \exp \left( -\rho U \sum_{\ell=0}^{k} \alpha_\ell \right) \right) + C_\theta \exp \left( -\frac{\rho U a}{1-\delta_2} (k+1)^{1-\delta_2} \right) + \frac{C_\zeta C_\rho}{a} \frac{1}{(k+1)^{\delta_2 - 2\delta_1}}. \tag{S211}
\]

From (S157) we have \( a = \left( \frac{C_a}{n^4} \right)^{\frac{1}{2}} \) and from (S161) we have \( C_\xi = a \bar{C}_\xi \), where

\[
\bar{C}_\xi = \left( \frac{2 C_\zeta L^4 C_w}{n^{\frac{4}{3}}} + \frac{2 C_\zeta L^2 \bar{L}^2 W_2}{n^{\frac{5}{3}}} + L^2 W_2 + \frac{2 C_\zeta \bar{L}^2 d_w}{n^{\frac{7}{3}}} \right). \tag{S212}
\]

Thus it follows from (S211) that

\[
F(p_{k+1}(\bar{w})) \leq \left( F(p_0(\bar{w})) + \sum_{\ell=0}^{k} C_\xi \exp \left( \frac{\rho U}{1-\delta_2} \left( \ell^{1-\delta_2} \right) \right) \exp \left( -\rho U \sum_{\ell=0}^{k} \alpha_\ell \right) \right) + C_\theta \exp \left( -\frac{\rho U a}{1-\delta_2} (k+1)^{1-\delta_2} \right) + \bar{C}_\xi C_\rho \frac{1}{(k+1)^{\delta_2 - 2\delta_1}}. \tag{S213}
\]

Considering the term \( L^2 W_2 \) in (S212), recall \( W_2 = \frac{2 \alpha (2d_w + n a \mu_a) (\delta_2 - \delta_1)}{b^2 \lambda_2 (L)^2 \delta_1} \exp \left( W_1 2^{1-\delta_1} \right) \) from Theorem \( \Box \) which, together with \( a \) in (S150), leads to

\[
W_2 = \frac{(\delta_2 - \delta_1) \exp \left( W_1 2^{1-\delta_1} \right)}{b^2 \lambda_2 (L)^2 \delta_1} \left( \frac{4 d_w}{n^{1/3}} \left( \frac{\rho U (3 \delta_2 - 1)}{25 \bar{L}^4 \delta_2} \right)^{\frac{1}{3}} + \frac{2 \mu_a}{n^{2/3}} \left( \frac{\rho U (3 \delta_2 - 1)}{25 \bar{L}^4 \delta_2} \right)^{\frac{2}{3}} \right) \tag{S214}
\]

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\[ \frac{(\delta_2 - \delta_1)}{n^{1/3}b^2\lambda_2(L)^2\delta_1} \exp \left( \frac{W_1}{25L^4\delta_2} \right)^{1/3} + \frac{2\mu_g}{n^{1/3}} \left( \frac{\rho_U(3\delta_2 - 1)}{25L^4\delta_2} \right)^{2/3} \]

\[ = \frac{\bar{C}_{w_2}}{n^{1/3}}. \]  

(S215)

where

\[ \bar{C}_{w_2} = \frac{(\delta_2 - \delta_1)}{b^2\lambda_2(L)^2\delta_1} \exp \left( \frac{W_1}{25L^4\delta_2} \right)^{1/3} + \frac{2\mu_g}{n^{1/3}} \left( \frac{\rho_U(3\delta_2 - 1)}{25L^4\delta_2} \right)^{2/3}, \]

(S216)

Substituting (S216) into (S212) yields

\[ \bar{C}_\xi = \left( \frac{2C_a^2L^4C_w}{n^{5/2}} + \frac{2C_a^2L^2L^2\bar{C}_{w_2}}{n^{1/3}} + \frac{L^2\bar{C}_{w_2}}{n^{1/3}} + \frac{2C_a^2L^2d_w}{n^{5/2}} \right). \]  

(S217)

Now substituting (S218) into (S213), we have

\[ F(p_{k+1}(\bar{w})) \leq F(p_{t_0}(\bar{w})) + \sum_{\ell=0}^{k} C_\xi \exp \left( \frac{\rho_Ua}{1-\delta_2} \left( \ell^{1-\delta_2} \right) \right) \exp \left( -\rho_U \sum_{\ell=0}^{k} \alpha_\ell \right) \]

\[ + \left( \frac{2C_a^2L^4C_wC_\rho}{n^{5/2}} + \frac{2C_a^2L^2L^2\bar{C}_{w_2}C_\rho}{n^{1/3}} + \frac{L^2\bar{C}_{w_2}C_\rho}{n^{1/3}} + \frac{2C_a^2L^2d_wC_\rho}{n^{5/2}} \right) \frac{1}{(k+1)^{\delta_2-2\delta_1}} \]

\[ + C_\theta \exp \left( -\frac{\rho_Ua}{1-\delta_2} (k+1)^{1-\delta_2} \right). \]  

(S219)

Note that the first and the third terms in (S219) are exponentially decaying while the second term is polynomial in \( k \). However the polynomial term decreases with the number of agents, \( n \). We further rewrite (S219) as

\[ F(p_{k+1}(\bar{w})) \leq \left( F(p_{t_0}(\bar{w})) + \bar{C}_{\xi_1} \right) \exp \left( -\rho_U \sum_{\ell=0}^{k} \alpha_\ell \right) + \frac{1}{n^{1/3}} \frac{\bar{C}_{\xi_2}}{(k+1)^{\delta_2-2\delta_1}} \]

\[ + \frac{\bar{C}_{\xi_3}}{\exp \left( \frac{\rho_Ua}{1-\delta_2} (k+1)^{1-\delta_2} \right)}. \]  

(S220)
where
\[
\bar{C}_{\ell_1} = \sum_{\ell=0}^{k} C_\xi \exp{(\rho_U a)} \exp\left(\frac{\alpha_{\rho_U}}{1 - \delta_2} \left(\ell^{1 - \delta_2}\right)\right), \quad \bar{k} = \left[\left(\frac{2 \delta_2 - 2 \delta_1}{\rho_U a}\right)^{1/\gamma_2}\right]
\]

\[
\bar{C}_{\ell_2} = \left(\frac{2 C_a^\frac{5}{3} L^4 C_w C_\rho}{n^{\frac{2}{3}}} + \frac{2 C_a^\frac{5}{3} L^2 L^2 C_w^2 C_\rho}{n^{\frac{2}{3}}} + L^2 C_w C_\rho + \frac{2 C_a^\frac{5}{3} L^2 d_w C_\rho}{n}\right)
\]

\[
\bar{C}_{\ell_3} = W_3(2a^3 L^2 L^2 + L^2 a) \exp(\rho_U a) \left(\sum_{\ell=0}^{\bar{\ell}} \exp\left(\frac{\rho_U a}{1 - \delta_2} (\ell + 1)^{1 - \delta_2} - W_1(\ell + 1)^{1 - \delta_1}\right) + \kappa \frac{\rho_U a}{1 - \delta_1} \Gamma \left(\frac{1}{1 - \delta_1}\right)\right)
\]

\[
\kappa = -\left(\frac{\rho_U a}{1 - \delta_2}(\overline{\ell} + 1)^{\delta_1 - \delta_2} - W_1\right), \quad \overline{\ell} = \left[\left(\frac{\rho_U a}{(1 - \delta_2)W_1}\right)^{\frac{1}{1 - \delta_1}} - 1\right]
\]

\[
C_\xi = a \left(\frac{2 C_a^\frac{5}{3} L^4 C_w}{n^{\frac{2}{3}}} + \frac{2 C_a^\frac{5}{3} L^2 L^2 W_2}{n^{\frac{2}{3}}} + L^2 W_2 + \frac{2 C_a^\frac{5}{3} L^2 d_w}{n^{\frac{2}{3}}}\right)
\]

\[
C_\rho = \frac{2(\delta_2 - \delta_1)}{\rho_U \delta_2} \exp\left(\frac{\rho_U a}{1 - \delta_2} 2^{1 - \delta_2}\right)
\]

\[
C_a = \frac{\rho_U (3\delta_2 - 1)}{25L^4 \delta_2} \Rightarrow a = \left(\frac{\rho_U (3\delta_2 - 1)}{25n^4 L^4 \delta_2}\right)^{\frac{1}{4}}
\]

\[
C_w = \max\left\{\mathbb{E}\left[\|\tilde{w}(t_0)\|_2^2\right], \frac{2c_1 + \frac{4}{\rho_U}(c_2 + d_1)}{1 - \frac{24n^4 L^4 \delta_2 a^3}{\rho_U (3\delta_2 - 1)}}\right\}
\]

\[
\bar{C}_{w_2} = \frac{(\delta_2 - \delta_1) \exp\left(W_1 2^{1 - \delta_1}\right)}{\bar{b}^2 \chi_2(C)^2 \delta_1} \left(4d_\omega \left(\frac{\rho_U (3\delta_2 - 1)}{25L^4 \delta_2}\right)^{\frac{1}{4}} + \frac{2\rho_U (3\delta_2 - 1)^{\frac{3}{4}}}{n^{1/3}} \left(\frac{\rho_U (3\delta_2 - 1)}{25L^4 \delta_2}\right)^{\frac{3}{4}}\right)
\]

while $W_1$, $W_2$ and $W_3$ are defined in $\text{(S222)}$, $\text{(S226)}$ and $\text{(S229)}$, respectively, and $c_1$, $c_2$ and $d_1$ are given in Lemma $\text{[S5]}$. This concludes the proof of Theorem $2$. 

$\blacksquare$
S4.1 Lemmas used in the proof of Theorem 2

Lemma S5. For $f_i(\tilde{w})$, $\tilde{f}_i(\tilde{w})$ and $\kappa(\tilde{w})$ defined in \[101\], \[102\] and \[108\], respectively, we have

$$\int \kappa(\tilde{w}) (\nabla \cdot [f_i(\tilde{w})]) \ d\tilde{w} = - \int \nabla \log \left( \frac{p_i(\tilde{w})}{p^*(\tilde{w})} \right) \tilde{f}_i(\tilde{w}) \ d\tilde{w},$$  \hspace{1cm} (S230)

and

$$\int \kappa(\tilde{w}) (\nabla \cdot [\tilde{f}_i(\tilde{w})]) \ d\tilde{w} = - \int \nabla \log \left( \frac{p_i(\tilde{w})}{p^*(\tilde{w})} \right) \tilde{f}_i(\tilde{w}) \ d\tilde{w}. \hspace{1cm} (S231)$$

Proof: This lemma is similar to \[19\] Lemma 10.4.1. Here we use the identity for $x \in \mathbb{R}^d$, $b(x) : \mathbb{R}^d \to \mathbb{R}$ and $a(x) : \mathbb{R}^d \to \mathbb{R}$:

$$\nabla_x \cdot [b(x)a(x)] = (\nabla_x b(x))^\top a(x) + b(x)(\nabla_x \cdot [a(x)]). \hspace{1cm} (S232)$$

Thus we have

$$\kappa(\tilde{w}) (\nabla \cdot f_i(\tilde{w})) = \nabla \cdot [\kappa(\tilde{w})f_i(\tilde{w})] - (\nabla \kappa(\tilde{w}))^\top f_i(\tilde{w}). \hspace{1cm} (S233)$$

Note that

$$\int \nabla \cdot [\kappa(\tilde{w})f_i(\tilde{w})] \ d\tilde{w} = \int \ldots \int \sum_{i=1}^{d_w} \frac{\partial}{\partial \tilde{w}_i} (\kappa(\tilde{w})f_i(\tilde{w})) \ d\tilde{w}_1 \ldots d\tilde{w}_{i-1} d\tilde{w}_i d\tilde{w}_{i+1} \ldots d\tilde{w}_{d_w} \hspace{1cm} (S234)$$

$$= \sum_{i=1}^{d_w} \int \ldots \int \frac{\partial}{\partial \tilde{w}_i} (\kappa(\tilde{w})f_i(\tilde{w})) \ d\tilde{w}_1 \ldots d\tilde{w}_{i-1} d\tilde{w}_i d\tilde{w}_{i+1} \ldots d\tilde{w}_{d_w} \hspace{1cm} (S235)$$

$$= \sum_{i=1}^{d_w} \int \ldots \int ( (\kappa(\tilde{w})f_i(\tilde{w})) \mid_{\tilde{w}_i = \pm \infty}^{\tilde{w}_i = \pm \infty}) \ d\tilde{w}_1 \ldots d\tilde{w}_{i-1} d\tilde{w}_{i+1} \ldots d\tilde{w}_{d_w} \hspace{1cm} (S236)$$

$$= 0. \hspace{1cm} (S237)$$

The last equality holds when $( (\kappa(\tilde{w})f_i(\tilde{w})) \mid_{\tilde{w}_i = \pm \infty}) = 0$ for all $i = 1, \ldots, d_w$, which is satisfied under the condition that $p_i(\tilde{w}) \to 0$ as $\tilde{w}_i \to \pm \infty$. The same technical condition has been assumed in the literature, see e.g., one of the assumptions in \[109\] Theorem 3.1 and the “sufficiently fast decay at infinity” condition in \[4\] Appendix A.1.

It then follows that

$$\int \kappa(\tilde{w}) \nabla \cdot [f_i(\tilde{w})] \ d\tilde{w} = - \int \nabla \kappa(\tilde{w})^\top f_i(\tilde{w}) \ d\tilde{w}. \hspace{1cm} (S238)$$

Similar argument can be used to prove \[231\].

Lemma S6. Let $\tilde{w}^*$ denotes samples from the target-distribution $p^*$, i.e., $\tilde{w}^* \sim p^*$
and \( \bar{w}^* \) satisfies
\[
\mathbb{E} [\| \bar{w}^* \|^2] \leq c_1. \tag{S239}
\]

Let \( \bar{w}(t) \) denote samples from the distribution \( p_t(\bar{w}) \), i.e., \( \bar{w}(t) \sim p_t(\bar{w}) \). Suppose that the KL-divergence between the initial distribution \( p_{t_0}(\bar{w}) \) and the target distribution \( \bar{p}^* \), denoted as \( F(p_{t_0}(\bar{w})) \) is bounded by \( c_2 \), i.e.,
\[
F(p_{t_0}(\bar{w})) \leq c_2. \tag{S240}
\]

Also, suppose that \( 2\delta_2 - 2\delta_1 > 1 \) and a \( \bar{w} \) in \[15\] is chosen such that \( \frac{24n^4L^4_1\delta_1}{\rho_U(3\delta_2 - 1)} < 1 \). Then there exists a \( C_{\bar{w}} > 0 \) such that for all \( k \geq 0 \),
\[
\mathbb{E} [\| \bar{w}(t_k) \|^2] \leq C_{\bar{w}}, \tag{S241}
\]

where
\[
C_{\bar{w}} = \max \left\{ \mathbb{E}_{\bar{w}(t_0) \sim p_{t_0}} [\| \bar{w}(t_0) \|^2], \frac{2c_1 + \frac{4}{\rho_U}(c_2 + d_1)}{1 - \frac{24n^4L_1^4\delta_1^2}{\rho_U(3\delta_2 - 1)}} \right\} \tag{S242}
\]

and \( d_1 \) is a positive constant satisfying
\[
d_1 \geq \sum_{k=0}^{\infty} (2\alpha_k^2L_1^2L_2^2 + L^2\alpha_k) \mathbb{E}_{p(t_k)} \| \tilde{w}(t_k) \|^2 + 2\alpha_k^2L_2^2d_{\bar{w}}. \tag{S243}
\]

**Proof:** We prove the boundedness of \( \mathbb{E}_{p_{t_n}(\bar{w})} [\| \bar{w}(t_k) \|^2] \) by induction. Assume that there exists a sufficiently large \( C_{\bar{w}} > 0 \) such that
\[
\mathbb{E}_{p_{t_n}(\bar{w})} [\| \bar{w}(t_n) \|^2] \leq C_{\bar{w}}, \quad \forall \ n \leq k. \tag{S244}
\]

We next show that
\[
\mathbb{E}_{p_{t_{k+1}}(\bar{w})} [\| \bar{w}(t_{k+1}) \|^2] \leq C_{\bar{w}}. \tag{S245}
\]

Following the proof of \[19\] Lemma 6, we couple \( \bar{w}^* \) optimally with \( \bar{w}(t) \sim p_t(\bar{w}) \), i.e., \( (\bar{w}(t), \bar{w}^*) \sim \gamma \in \Gamma_{\text{opt}}(p_t(\bar{w}), \bar{p}^*) \). We then obtain
\[
\mathbb{E}_{\bar{w}(t_{k+1}) \sim p_{t_{k+1}}} [\| \bar{w}(t_{k+1}) \|^2] = \mathbb{E}[(\bar{w}(t_{k+1}), \bar{w}^*), \sim \gamma] [\| \bar{w}^* + \bar{w}(t_{k+1}) - \bar{w}^* \|^2]
\]
\[
\leq 2\mathbb{E} \bar{w}^* \sim \gamma [\| \bar{w}^* \|^2] + 2\mathbb{E} (\bar{w}(t_{k+1}), \bar{w}^*) \sim \gamma \| \bar{w}(t_{k+1}) - \bar{w}^* \|^2
\]
\[
\leq 2c_1 + 2\mathcal{W}_2^2(p_{t_{k+1}}(\bar{w}), \bar{p}^*) \tag{S246}
\]
\[
\leq 2c_1 + \frac{4}{\rho_U} F(p_{t_{k+1}}(\bar{w})), \tag{S247}
\]

where \( \mathcal{W}_2(\cdot, \cdot) \) denotes the Wasserstein metric between two distributions and the last inequality holds due to \[5\] Theorem 1.

From our analysis in \[149\], we have
\[
F(p_{t_{k+1}}(\bar{w})) \leq \exp(-\rho_U \alpha_k) F(p_{t_k}(\bar{w})) + \left(2\alpha_k^2L_1^4 \mathbb{E} \bar{w}(t_{k+1}) \| \bar{w}(t_k) \|^2 \right) \tag{S248}
\]

\[
\leq 2c_1 + \frac{4}{\rho_U} F(p_{t_{k+1}}(\bar{w})).
\]
We rewrite (S251) as

\[ F \] from which we further obtain

\[ \text{from (S252), we note that} \]

\[ g_k = (2\alpha_k^3 L^2 L^2 + L^2 \alpha_k) E_{p(\tilde{w}(t_k))} \| \tilde{w}(t_k) \|_{\ell_2}^2 + 2\alpha_k^2 L^2 d_w. \] (S252)

We rewrite (S251) as

\[ F(p_{t+1}(\tilde{w})) \leq \exp (-\rho U(\alpha_j)) F(p_{t_j}(\tilde{w})) + g_j + 2\alpha_j^3 \bar{L}^4 C_w, \quad \forall j \leq k, \] (S253)

from which we further obtain

\[ F(p_{t_{k+1}}(\tilde{w})) \leq \exp (-\rho U(\alpha_k)) F(p_{t_k}(\tilde{w})) + g_k + 2\alpha_k^3 \bar{L}^4 C_w \] (S254)

\[ \leq \exp (-\rho U(\alpha_k)) \exp (-\rho U(\alpha_{k-1})) F(p_{t_{k-1}}(\tilde{w})) + g_{k-1} + 2\alpha_{k-1}^3 \bar{L}^4 C_w \] (S255)

\[ \leq \exp (-\rho U(\alpha_{k-1})) F(p_{t_{k-1}}(\tilde{w})) + g_{k-1} + 2\alpha_{k-1}^3 \bar{L}^4 C_w \] (S256)

\[ \leq \exp \left( -\rho U \sum_{i=0}^{k} \alpha_{i} \right) F(p_{t_0}) + \sum_{i=0}^{k} g_i + 2C_w n^4 L^4 \sum_{i=0}^{k} \alpha_{i}^3 \] (S257)

\[ \leq F(p_{t_0}) + \sum_{i=0}^{\infty} g_i + 2C_w n^4 L^4 \sum_{i=0}^{k} \alpha_{i}^3 \] (S258)

\[ \leq c_2 + \sum_{i=0}^{\infty} g_i + 2C_w n^4 L^4 \sum_{i=0}^{k} \alpha_{i}^3. \] (S259)

Here we used the relation $\bar{L} \leq nL$. Since $\alpha_k^3 = \frac{a_k^3}{(k+1)^{3\delta_2}}$ and $2\delta_2 > 1$, it follows from (S252) that

\[ \sum_{i=0}^{k} \alpha_{i}^3 \leq a^3 + \int_{0}^{\infty} \frac{a^3}{(t+1)^{3\delta_2}} \, dt = a^3 + \frac{a^3}{3\delta_2 - 1} = \frac{3\delta_2}{3\delta_2 - 1} a^3. \] (S260)

From (S252), we note that $g_k \sim O \left( \frac{1}{(k+1)^{3\delta_2 - 1}} \right)$ because $E_{p(\tilde{w}(t_k))} \| \tilde{w}(t_k) \|_{\ell_2}^2 \sim O \left( \frac{1}{(k+1)^{3\delta_2 - 1}} \right)$ from (S2). Since $2\delta_2 - 2\delta_1 > 1$, $g_k$ is a summable sequence, that is, there exists a $d_1 > 0$ such that

\[ \sum_{i=0}^{\infty} g_i \leq d_1. \] (S261)

We obtain from (S250), (S260) and (S261) that

\[ F(p_{t_{k+1}}) \leq c_2 + d_1 + \frac{6\delta_2 a^3 n^4 L^4}{3\delta_2 - 1} C_w \] (S262)
which together with (S249) leads to

$$\mathbb{E}_{w(t_k+1) \sim p_{k+1}} \left[ \| \bar{w}(t_k+1) \|^2_2 \right] \leq 2c_1 + \frac{4}{\rho U} c_2 + \frac{24n^4L^4\delta_2 a_3}{\rho U(3\delta_2 - 1)} C_{\bar{w}} + \frac{4}{\rho U} d_1. \quad (S263)$$

Since $a$ is chosen such that

$$\frac{24n^4L^4\delta_2 a_3}{\rho U(3\delta_2 - 1)} < 1, \quad (S264)$$

it follows that

$$\mathbb{E}_{w(t_k+1) \sim p_{k+1}} \left[ \| \bar{w}(t_k+1) \|^2_2 \right] \leq 2c_1 + \frac{4}{\rho U} c_2 + \frac{24n^4L^4\delta_2 a_3}{\rho U(3\delta_2 - 1)} C_{\bar{w}} + \frac{4}{\rho U} d_1 \leq C_{\bar{w}} \quad (S265)$$

for any $C_{\bar{w}}$ such that

$$C_{\bar{w}} \geq \frac{2c_1 + \frac{4}{\rho U} (c_2 + d_1)}{1 - \frac{24n^4L^4\delta_2 a_3}{\rho U(3\delta_2 - 1)}}. \quad (S266)$$

One choice of $C_{\bar{w}}$ is

$$C_{\bar{w}} = \max \left\{ \mathbb{E}_{w(t_0) \sim p_{t_0}} \left[ \| \bar{w}(t_0) \|^2_2 \right], \frac{2c_1 + \frac{4}{\rho U} (c_2 + d_1)}{1 - \frac{24n^4L^4\delta_2 a_3}{\rho U(3\delta_2 - 1)}} \right\}. \quad (S267)$$

\section*{S5 Proof of Corollary 1}

From (S22b) we have

$$F(p_{t_k+1}(\bar{w})) \leq (F(p_{t_0}(\bar{w})) + \tilde{C}_{f_1}) \exp \left( -\rho_U \sum_{\ell=0}^{k} \alpha_\ell \right) + \frac{1}{n^{1/3} (k+1)^{\delta_2 - 2\delta_1}} \frac{\tilde{C}_{f_2}}{\exp \left( \frac{\rho U}{1-\delta_2} (k + 1)^{1-\delta_2} \right)}. \quad (S268)$$

When $\delta_2 \in (0.5, 1)$, from Lemma (S3) we have

$$\sum_{\ell=0}^{k} \alpha_\ell \geq \int_{0}^{k} \frac{a}{(x + 1)^{\delta_2}} dx = \frac{a(k+1)^{1-\delta_2}}{1-\delta_2} - \frac{a}{1-\delta_2}. \quad (S269)$$

Thus

$$\exp \left( -\rho_U \sum_{\ell=0}^{k} \alpha_\ell \right) \leq \exp \left( \frac{a\rho U}{1-\delta_2} \right) \exp \left( -\frac{a\rho U}{1-\delta_2} (k + 1)^{1-\delta_2} \right). \quad (S270)$$

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and (S268) can be written as
\[
F(p_t k + 1 (\bar{w})) \leq (F(p_t_0 (\bar{w})) + C_{r_1}) \exp \left( \frac{a \rho U}{1 - \delta_2} \right) \exp \left( -\frac{a \rho U}{1 - \delta_2} (k + 1)^{1 - \delta_2} \right) + \frac{1}{n^{1/3}} \frac{C_{r_2}}{(k + 1)^{\delta_2 - 2\delta_1}} + C_{r_3} \exp \left( -\frac{a \rho U}{1 - \delta_2} (k + 1)^{1 - \delta_2} \right)
\]  
(S271)

Now define constants \(Q_1\) and \(Q_2\) as
\[
Q_1 = (F(p_t_0 (\bar{w})) + C_{r_1}) \exp \left( \frac{a \rho U}{1 - \delta_2} \right) + C_{r_3}
\]  
\(Q_2 = \frac{C_{r_2}}{n^{1/3}}.
\]  
(S272)

Thus we have
\[
F(p_t k + 1 (\bar{w})) \leq Q_1 \exp \left( -\frac{a \rho U}{1 - \delta_2} (k + 1)^{1 - \delta_2} \right) + Q_2 \frac{1}{(k + 1)^{\delta_2 - 2\delta_1}}
\]  
(S273)

We would like to find a \(k\) such that
\[
F(p_t k + 1 (\bar{w})) \leq \epsilon,
\]  
(S274)

which is satisfied if
\[
Q_1 \exp \left( -\frac{a \rho U}{1 - \delta_2} (k + 1)^{1 - \delta_2} \right) \leq \frac{\epsilon}{2}
\]  
(S275)

and
\[
Q_2 \frac{1}{(k + 1)^{\delta_2 - 2\delta_1}} \leq \frac{\epsilon}{2}.
\]  
(S276)

From (S276) we have
\[
k \geq \left( \frac{1 - \delta_2}{a \rho U} \log \left( \frac{2Q_1}{\epsilon} \right) \right)^{\frac{1}{1 - \delta_2}}
\]  
(S277)

and from (S277) we have
\[
k \geq \left( \frac{2Q_2}{\epsilon} \right)^{\frac{1}{\delta_2 - 2\delta_1}}
\]  
(S278)

Therefore, for all \(k \geq k^*\), we have
\[
F(p_t k (\bar{w})) \leq \epsilon,
\]  
(S279)

where
\[
k^* = \max \left\{ \left( \frac{1 - \delta_2}{a \rho U} \log \left( \frac{2Q_1}{\epsilon} \right) \right)^{\frac{1}{1 - \delta_2}}, \left( \frac{2Q_2}{\epsilon} \right)^{\frac{1}{\delta_2 - 2\delta_1}} \right\}
\]  
(S280)
Table S1: Test accuracy (%) for different approaches after 10 epochs

|        | SGD | C-ULA | Agent 1 | Agent 2 | Agent 3 | Agent 4 | Agent 5 |
|--------|-----|-------|---------|---------|---------|---------|---------|
| MNIST  | 98.15 | 98.16 | 98.52   | 98.52   | 98.39   | 98.45   | 98.47   |

This concludes the proof of Corollary 1.

S6  Numerical Experiments

S6.1 Parameter estimation for Gaussian mixture

For the centralized setting, 100 data samples are drawn from the mixture of Gaussians in Section 5.1. For D-ULA, these 100 samples were randomly divided into 5 data sets of 20 samples, one for each of the five agents in the network. Both C-ULA and D-ULA are run for 1000000 epochs using their respective batch gradients. Step-size $\alpha_k = \alpha_0 / (b_1 + k)^{\delta_2}$ is varied from 0.01 to 0.0001 similar to [26] with consensus step-size, $\beta_k = \beta_0 / (b_2 + k)^{\delta_1}$ in the interval [0.36, 0.24]. Fig. S1 shows estimated posteriors from C-ULA and the proposed approach. Posteriors estimated by the D-ULA replicate the true posterior with samples from both the modes.

S6.2 Bayesian logistic regression

Expressions for time-varying step-size $\alpha_k$ and $\beta_k$ are same as in Section S6.1 with $\alpha_0 = 0.004$, $b_1 = 230$, $\delta_2 = 0.55$ for C-ULA and $\alpha_0 = 0.00082$, $b_1 = 230$, $\delta_2 = 0.55$, $\beta_0 = 0.48$ $b_2 = 230$, $\delta_1 = 0.05$ for D-ULA. Data is processed in batches of 10 for both approaches with 10 epochs through the whole data set for 50 runs. Accuracy at each iteration averaged over 50 runs for C-ULA and the 5 agents in D-ULA is shown in Fig. S2. The shaded region of the figure indicates 1 standard deviation. Zoomed version of the accuracy with centralized ULA shown in Fig. S6.2.b indicates a faster convergence of D-ULA to 84.38% in 1040 iterations when compared to C-ULA which converges to the final accuracy of 83.89%.

S6.3 Decentralized Bayesian learning for handwritten digit recognition

Variations in step-size, $\alpha_k$ and $\beta_k$ are similar to Section S6.1 with $\alpha_0 = 0.00024$, $b_1 = 230$, $\delta_2 = 0.55$ for stochastic gradient descent (SGD), $\alpha_0 = 0.0003$, $b_1 = 230$, $\delta_2 = 0.55$ for C-ULA, and $\alpha_0 = 0.00032$, $b_1 = 230$, $\delta_2 = 0.55$, $\beta_0 = 0.48$, $b_2 = 230$, $\delta_1 = 0.05$ for D-ULA. Data sets for training are processed in batches of 1024, 1024, and 256 images for SGD, C-ULA, and D-ULA, respectively. Table S1 summarizes test accuracy after 10 epochs for SGD, C-ULA, and D-ULA. Fig. S3 shows prediction probability density for MNIST and SVHN data sets using all the approaches considered.
Figure S1: True and estimated posteriors
Figure S2: Accuracy on test sets averaged over 50 runs
Figure S3: Probability of predicted labels