Species Distribution Model Predictions of the Critically Endangered Grey Nurse Shark in Australia

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Research Article

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Species distribution model predictions of the Critically Endangered grey nurse shark in Australia

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Abstract

Species distribution models (SDMs) are commonly used to forecast how threatened species are influenced by climate change. The grey nurse shark (*Carcharias taurus*) is a critically endangered species inhabiting both the east and west coasts of Australia, with negligible genetic interchange between the two populations. I used Generalized Linear Models (GLM), Maximum Entropy (MaxEnt) models and Boosted Regression Trees (BRT) to predict the distribution of the grey nurse shark. The data were a sample of presence-only data, derived from the known grey nurse shark sighting locations, from the east coasts of Australia, with pseudo-absences generated and bootstrapped from a restricted background. I verified these models using leave-one-out cross validation and model metrics including AICc, BIC, percentage of deviance explained, leave-one-out cross-validated $R^2$, AUC, maximum Cohen’s Kappa, specificity and sensitivity. Cross-validated $R^2$ was used as an overall comparison method across model types. I performed out-of-source validation by comparing model projection with the distributional range of the ragged tooth shark (*Carcharias taurus*) in South Africa. The prediction of the selected model was consistent with the current distributional range of the ragged tooth shark.
**Introduction**

Species distribution modelling has been widely used in conservation biology and ecology \(^1\) for discovering new species \(^2\), prediction of distributional ranges \(^3,4\), modelling invasive species’ potential distribution \(^5\), and detecting climate impacts on the habitat suitability \(^6\). The species distribution model has been used in many domains, in combination with ecological niche theory, at both individual and community levels \(^7-10\).

These distribution models use different mechanisms to infer the characteristics of the preferred habitat of target species \(^11\). Many methods have been subjected to model comparison and performance evaluation, with no obviously best method that is suitable for all situations \(^12\). Model accuracy is difficult to quantify because the evaluation methods (threshold-dependent and threshold-independent criteria) are susceptible and there is typically a lack of sufficient data to do independent, out-of-sample performance tests \(^13,14\).

The selection of pseudo-absence data can also exert a strong influence on subsequent predictions \(^15\). For instance, the questions of how large a background region to cover, and how many pseudo-absence points are optimal, are not settled issues. Some studies have shown that the relative importance of predictors and performance statistics such are sensitive to scale and background \(^16,17\). However, the relationship between species response curve fitting and pseudo-absence data selection remains an inadequately resolved problem \(^8\).
As an alternative to making assumptions about pseudo-absences, presence-only models such as MaxEnt\(^{18}\), GARP\(^{19}\), and ENFA\(^{20}\) can be used when no absence data is available. They detect the marginality of the species in the entire background\(^{21}\). Again, however, the question of how background scale and frequency distribution influence the model performance to the entire environment is not well studied.

I developed a case study of the grey nurse shark (*Carcharias taurus*, Rafinesque 1810), a coastal marine predatory fish with a small presence sample size and very wide distribution along the coast of Australia, to test the abiotic factor impact on this top marine predator’s distribution. In Australia, several broad-scale surveys in the states of New South Wales, Queensland and Western Australia have been done to identify the species’ aggregation sites and migration patterns; however, its west-coast distribution is poorly quantified and only one aggregation site was identified\(^{22}\). Historically, the east and west coast populations have been separated and have negligible genetic interchange\(^{23}\). It has been argued that the model with better transferability is better algorithm\(^{24}\). As such, I used the east coast sighting locations for model fitting and projected the best fitted model in South African coast for model testing. Presence-absence and Presence-only models GLM, BRT and MaxEnt were fitted.

**Methods**
Species data

I collated data describing the major sighting sites of the east coast population of grey nurse sharks from the *Issues Paper for the Grey Nurse Shark (Carcharias taurus)* generated by Australian Government (https://environment.gov.au/system/files/resources/91e141d0-47aa-48c5-8a0f-992b9df960fe/files/issues-paper-grey-nurse-shark-carcharias-taurus.pdf). The original presence data consisted of 153 points. However, there were 17 points falling on the grids where the elevation values were positive. I deleted these data points from the presence data. There was one presence point lying on the grid where the bathymetric slope value was zero. I removed this point in order to apply logarithm transformation to the bathymetric slope data. Thus, the number of the input presence points was 135.

Environmental data

The four prime predictors were bathymetric slope, depth, 15-year January Sea Surface Temperature (SST) mean and 15-year January interannual SST variance. The Australian Bathymetry and Topography Grid dataset was used for model fitting. The resolution of this dataset is 0.0025°. The ETOPO1 1 Arc-Minute Global Relief Model was obtained for model projection in South African coastal area. DEM Surface Tools in ArcGIS 10 (ESRI 2011. ArcGIS Desktop: Release 10. Redlands, California, USA) was used to calculate the bathymetric slope data set. The SST mean and variance
predictors were generated from the 4 km Advanced Very High Resolution Radiometer (AVHRR) Pathfinder Version 5.0. The SST mean and variance were bilinearly interpolated, so that the resolutions of the four predictors were the same.

Logarithm transformation was applied to bathymetric slope and January SST mean. The square of the log-transformed January SST mean was calculated for GLM and BRT. I scaled bathymetric depth and interannual SST variance according to the following equation (1):

\[ X_{ij}^n = \frac{X_{ij} - X_{j,\text{min}}}{X_{j,\text{max}} - X_{j,\text{min}}} \]  

where \( X_{ij}^n \) and \( X_{ij} \) are the scaled and non-scaled \( j \)th variable values of point \( i \). \( X_{j,\text{max}} \) and \( X_{j,\text{min}} \) are the maximum and minimum values of \( j \)th variable. Because MaxEnt applies this scaling to all its predictors automatically, the scaled SST variance, log-transformed bathymetric slope, non-transformed SST mean and non-scaled bathymetric depth were used as input for MaxEnt models.

Model types

I employed three prevalent species distribution models: GLM 25, MaxEnt 18 and BRT 26,27. Generalized linear models are widely used in species distribution modelling 28,29. GLM is generalized multiple variates regression analysis using the link function 30. I used binomial distribution with a logit link function to fit the presence and pseudo-absence dataset. MaxEnt (version 3.3.3k, AT&T Labs Research) is a presence-only modelling technique using both background and presence data. MaxEnt fits
coefficients by minimizing the regularized negative log-likelihood to achieve the maximum entropy of the predicted distribution. I used linear and quadratic features for MaxEnt model fitting. Regression Tree model makes use of least-square regression to estimate parameters. The interaction depth was set to 3 to avoid overfitting, and the bag fraction was set to 1. The shrinkage value and the number of trees were set to 0.01 and 200, respectively.

Pseudo-absence point generation

Grey nurse sharks are mainly found in the shallow neritic waters, so a common sample bias is always detected in the data. I used the restricted background method to reduce this sample bias. I restricted the maximum depth of background layers to 93 m, for the maximum depth of the presence points was 89 m. I extended the range of the background layers to 1° further than that of the presence points (see Figure 1). I also eliminated the background data points lying on the grids where the bathymetric slope values were zero. The number of pseudo-absence points in each data set was set to 1350, which was 10 times as big as the number of presence points.

I bootstrapped the pseudo-absences from the restricted background using randomly sampling method with replacement in R statistical software (R Core Team). I generated 500 bootstrap pseudo-absence samples to form 500 data sets. Each data set had 135 presences and 1350 pseudo-absences. I then randomly split each data set
Model evaluation and out-of-source validation

Various model metrics were used for model evaluation. Firstly, the sample size corrected Akaike information criterion (Arc) and Bayesian Information Criterion (BIC) $^{36}$ were calculated for GLM and MaxEnt models $^{37}$. Secondly, leave-one-out cross-validated coefficient of determination (cross-validated $R^2$) $^{38}$ were computed for measuring the goodness-of-fit of the candidate models. Cross-validated $R^2$ was used as a universal model evaluation method across different model types. For MaxEnt, the logistic output values were used to calculate the cross-validated $R^2$. Thirdly, threshold-independent Area Under the Curve (AUC) $^{39}$, maximum Cohen’s Kappa $^{40}$, specificity and sensitivity $^{37,39,41}$ were also calculated by leave-one-out cross validation. I then estimated the median, 5$^{th}$ and 95$^{th}$ percentiles for each model metric $^{42}$.

I projected the median cross-validated $R^2$ selected candidate model in the coastal area of South Africa as an out-of-source validation. In this validation, the bathymetric depth was restricted and the points where bathymetric slope values were zero were eliminated. All the variables in the out-of-source dataset were transformed and scaled according to the methods mentioned before.
Results

Model evaluation

I used cross-validated $R^2$, which is asymptotically equivalent to AIC-type criteria, as a universal method for model selection across the three model types. The model ranking results were shown in Table 1, Table 2 and Table 3.

According to the median values of these model metrics, one of the GLM candidate models gained the highest median value of cross-validated $R^2$.

This model comprised bathymetric slope, depth, January SST mean and January SST mean square as predictors. The frequencies that this model was selected among GLM candidate models by AICc, BIC and cross-validated $R^2$ were 72.7%, 99.8% and 77.6%. The median value of cross-validated $R^2$ of this model was 0.3569, which was the highest across three model types. The median values of AUC, Maximum Cohen’s Kappa and the sum of Sensitivity and Specificity of this model were also the highest across three model types.

With respect to MaxEnt models, the cross-validated $R^2$ of some candidate models were negative. The ranking results by BIC and cross-validated $R^2$ for MaxEnt candidate models were the same. The MaxEnt candidate model selected by AICc was the full model while BIC and cross-validated $R^2$ selected one was comprised of bathymetric slope, depth, January SST mean and the square of bathymetric slope, the square of bathymetric depth and the square of January SST mean as predictors. The frequencies that the full model was selected as the best model by AICc, BIC and
cross-validated $R^2$ were 59.8%, 19.7% and 39% while as for the BIC selected model, the frequencies were 40.2%, 78.9% and 61%. The Sensitivity values of MaxEnt models were higher than those of GLM and BRT models.

The cross-validated $R^2$ selected BRT candidate model consisted of bathymetric slope, depth, January SST mean and January SST mean square as dependent variables. The median value of cross-validated $R^2$ for this model was 0.3058. The median value of AUC gained by this model was 0.9094, which was lower than that of the best GLM model.

Out-of-source validation

I projected the GLM candidate model with the highest median value of cross-validated $R^2$ in the South African and Australian coastal areas. Model projections were consistent with the current distributional ranges of the ragged tooth shark (*Carcharias taurus*) and the grey nurse shark (*Carcharias taurus*). Figure 2 and Figure 3 were averaged predictions from 500 simulations. The distributional range of ragged tooth sharks in the coast of South Africa was from around Sodwana Bay, Kwazulu-Natal to Cape Town. As for the east coast population in Australia, the sighting locations of the grey nurse shark ranged from Fraser Island, Queensland to Gabo Island, Victoria (*Issue paper for the grey nurse shark (Carcharias taurus)*). It is reported in the *Issue paper* that the range of the west coast population in Australia was from Exmouth, Western Australia to the coastal area near Cocklebiddy in the
Great Australian Bight. The model predictions were in line with these ranges (see Figure 2 and Figure 3).

**Discussion**

The predicted habitat suitability was mainly dependent on the multivariate response surface. Carefully examining the multivariate response surface is important. The response curve of January SST mean of the cross-validated $R^2$ selected GLM model was a narrow bell-shaped curve, which may resemble the realised niche of the species. Grey nurse sharks’ aggregation sites are always found around inshore islands and rocky reefs with pinnacles, gutters and rocky caves. Thus, the bathymetric slope might be used as a surrogate for depicting the complex sea bed structure that the shark prefers. As for the bathymetric depth, the model indicated a trend that shallow waters were more preferable than deep waters for the species. However, although the predicted habitat suitability in waters deeper than 90 m was close to zero, the maximum depth to which the shark dives is deeper than 200 m, indicating a sampling bias influenced the model performance.

Various model evaluation methods were explored in this study. Overall, although the median value of cross-validated $R^2$ of the best fitted GLM model was not very high, the model’s projections were consistent with the current distributional ranges of this species in South Africa and Australia, indicating that it captured the inherent responses to the environmental gradients. From the model selection results, it was
easy to observe that the SST mean, bathymetric slope and depth were dominant
variables. In this study, the interannual SST variance was highly correlated with the
SST mean, which may result in model selection bias.

Data availability

The datasets generated and analysed during the current study are available from the
corresponding author on reasonable request.

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**Author contributions**

G. S. designed the case study, collected the data, and did the analysis. The first draft of the manuscript was written and then revised by G. S.
Competing interests

The author declares no competing interests.
Fig 1 Sighting locations and restricted background in the east coast of Australia
Fig 2 The projection of the best GLM candidate model in South African coastal area (the resolution of environmental layers was interpolated into 0.04394°)
Fig 3 The projection of the best GLM candidate model in the coasts of Australia (the resolution of environmental layers was aggregated from 0.0025° to 0.1°)
Table 1: Median values of model metrics (Slope.l: log-transformed bottom slope; JanSSTm.l: log-transformed January SST mean; JanSSTm.var: scaled interannual variance of January SST; Depth.c: scaled bathymetric depth; JanSSTm: January SST mean; Depth: bathymetric depth; LogL: log-likelihood; AICc: a sample size corrected AIC; dAICc: the difference between AIC value of a candidate model and the smallest AIC value in the set of candidate models; wAICc: AICc weight; BIC: Bayesian Information Criterion; dBIC: the difference between BIC value of a candidate model and the smallest BIC value in the set of candidate models; wBIC: BIC weight; pcdev: percentage of explained deviance; cvweight: CV weight; r2n.pred: cross-validated R2)

| GLM (median) | k | LogL | AICc | BIC | dAICc | wAICc | dBIC | wBIC | pcdev | r2n_pred | AUC | Kappa | specificity | sensitivity |
|--------------|---|------|------|-----|-------|-------|------|-------|-------|--------|-----|-------|-------------|-------------|
| PA ~ Slope.l+JanSSTm.l+Depth.c+JanSSTm2^2 | 5 | -256.3375 | 522.7158 | 549.1536 | 0.0000 | 0.6168 | 0.0000 | 0.9554 | 42.9135 | 0.3569 | 0.9257 | 0.5330 | 0.8467 | 0.8741 |
| PA ~ Slope.l+JanSSTm.l+Depth.c+JanSSTm.var.c+JanSSTm2^2 | 6 | -255.6300 | 523.3172 | 555.0343 | 0.9521 | 0.3832 | 6.2314 | 0.0424 | 43.0711 | 0.3557 | 0.9254 | 0.5318 | 0.8474 | 0.8667 |
| PA ~ Slope.l+Depth.c | 3 | -273.0408 | 552.0979 | 567.9688 | 29.5815 | 0.0000 | 18.4995 | 0.0001 | 39.1937 | 0.3247 | 0.9136 | 0.4972 | 0.8341 | 0.8741 |
| PA ~ Slope.l+JanSSTm.var.c | 3 | -333.1319 | 672.2801 | 688.1510 | 150.2661 | 0.0000 | 139.2166 | 0.0000 | 25.8114 | 0.1904 | 0.8584 | 0.3808 | 0.7670 | 0.8148 |
| PA ~ Slope.l | 2 | -336.7513 | 677.5107 | 688.0940 | 155.1608 | 0.0000 | 138.9040 | 0.0000 | 25.0503 | 0.1842 | 0.8563 | 0.3778 | 0.7593 | 0.8000 |
| PA ~ Depth.c | 2 | -353.3590 | 710.7261 | 721.3094 | 188.4702 | 0.0000 | 172.2746 | 0.0000 | 21.3068 | 0.1521 | 0.8314 | 0.3504 | 0.7022 | 0.8000 |
| PA ~ JanSSTm.var.c+JanSSTm.l+JanSSTm2^2 | 4 | -372.3683 | 752.7639 | 773.9196 | 230.2243 | 0.0000 | 224.6572 | 0.0000 | 17.0734 | 0.1116 | 0.8052 | 0.2957 | 0.7085 | 0.7407 |
| MaxEnt (median) | | | | | | | | | | | | | | |
| PA ~ Slope.l+JanSSTm+Depth+JanSSTm.var.c+Slope.l^2+JanSSTm2^2 | 8 | -714.7772 | 1445.9443 | 1485.3462 | 0.0000 | 0.6864 | 6.8351 | 0.0316 | 25.9188 | 0.3124 | 0.9209 | 0.5185 | 0.7919 | 0.9037 |
| Depth2^2+JanSSTm.var.c^2 | | | | | | | | | | | | | | |
| PA ~ Slope.l+JanSSTm+Depth+Slope.l+JanSSTm2^2+Depth2^2 | 6 | -718.7651 | 1449.9570 | 1480.5641 | 1.5669 | 0.3136 | 0.0000 | 0.9611 | 25.5055 | 0.3132 | 0.9213 | 0.5156 | 0.7807 | 0.9185 |
| PA ~ Slope.l+Depth2^2 | 3 | -738.8469 | 1485.8564 | 1506.2684 | 40.4540 | 0.0000 | 27.3045 | 0.0000 | 23.4242 | 0.2743 | 0.9125 | 0.4957 | 0.7630 | 0.9185 |
| PA ~ Slope.l+JanSSTm.var.c+Slope.l+JanSSTm.var.c^2 | 4 | -825.7029 | 1659.0461 | 1678.1177 | 214.0562 | 0.0000 | 199.4114 | 0.0000 | 14.4222 | -0.0861 | 0.8564 | 0.3832 | 0.4681 | 0.9630 |
| PA ~ Slope.l+Slope.l^2 | 2 | -830.9622 | 1665.9493 | 1676.4112 | 221.0436 | 0.0000 | 197.2694 | 0.0000 | 13.8772 | -0.1201 | 0.8525 | 0.3795 | 0.4570 | 0.9630 |
| PA ~ Depth1+Depth2^2 | 2 | -854.2826 | 1712.6496 | 1723.1434 | 266.8083 | 0.0000 | 243.7734 | 0.0000 | 11.4602 | -0.3048 | 0.8306 | 0.3525 | 0.4222 | 0.9704 |
| PA ~ JanSSTm.var.c+JanSSTm+JanSSTm2+JanSSTm2^2 | 4 | -876.5274 | 1761.3649 | 1782.2378 | 316.2833 | 0.0000 | 302.9687 | 0.0000 | 9.1547 | -0.6422 | 0.8075 | 0.2915 | 0.1237 | 1.0000 |
| BRT (median) | | | | | | | | | | | deviation | | | |
| PA ~ Slope.l+JanSSTm+l+Depth.c+JanSSTm2 | - | 526.6876 | - | - | - | - | - | - | 41.3532 | 0.3058 | 0.9094 | 0.4958 | 0.8644 | 0.7926 |
| PA ~ Slope.l+JanSSTm+l+Depth.c+JanSSTm2 | - | 526.4892 | - | - | - | - | - | - | 41.3753 | 0.3047 | 0.9091 | 0.4947 | 0.8644 | 0.7926 |
| PA ~ Slope.l+Depth.c | - | 542.9465 | - | - | - | - | - | - | 39.5428 | 0.2986 | 0.8994 | 0.4945 | 0.8622 | 0.7852 |
| PA ~ Slope.l+JanSSTm.var.c | - | 644.0666 | - | - | - | - | - | - | 28.2830 | 0.1764 | 0.8223 | 0.3818 | 0.7889 | 0.7556 |
| PA ~ Slope.l | - | 657.9448 | - | - | - | - | - | - | 26.7377 | 0.1714 | 0.8129 | 0.3743 | 0.7859 | 0.7630 |
| PA ~ Depth.c | - | 695.0521 | - | - | - | - | - | - | 22.6088 | 0.1430 | 0.7986 | 0.3489 | 0.6907 | 0.7704 |
| PA ~ JanSSTm.var.c+JanSSTm+l+JanSSTm2 | - | 708.2194 | - | - | - | - | - | - | 21.1396 | 0.0943 | 0.7763 | 0.2684 | 0.6844 | 0.7778 |
Table 2: 5th percentiles of model metrics (Slope.l: log transformed bottom slope; JanSSTm.l: log-transformed January SST mean; JanSSTm.l: the square of the log-transformed January SST mean; JanSSTvar.c: scaled interannual variance of January SST; Depth.c: scaled bathymetric depth; JanSSTm: January SST mean; Depth: bathymetric depth; LogL: log-likelihood; AICc: a sample size corrected AIC; dAICc: the difference between AIC value of a candidate model and the smallest AIC value in the set of candidate models; wAICc: AICc weight; BIC: Bayesian Information Criterion; dBIC: the difference between BIC value of a candidate model and the smallest BIC value in the set of candidate models; wBIC: BIC weight; BIC differences; pcdev: percentage of explained deviance; cvweight: CV weight; r2n_pred: cross-validated R^2)

| GLM | k | LogL | AICc | BIC | dAICc | wAICc | dBIC | wBIC | pcdev | r2n_pred | AUC | Kappa | specificity | sensitivity |
|-----|---|------|------|-----|-------|-------|------|------|-------|----------|-----|-------|--------------|------------|
| PA ~ Slope.l+JanSSTm.l+Depth.c+JanSSTm.l^2 | 5 | -267.2309 | 496.6940 | 523.1318 | 0.0000 | 0.2235 | 0.0000 | 0.7942 | 40.4875 | 0.3281 | 0.9191 | 0.5006 | 0.8341 | 0.8593 |
| PA ~ Slope.l+JanSSTm.l+Depth.c+JanSSTvar.c+JanSSTm.l^2 | 6 | -266.3031 | 497.1890 | 528.9062 | 0.0000 | 0.2705 | 2.7890 | 0.0257 | 40.6942 | 0.3268 | 0.9188 | 0.5009 | 0.8348 | 0.8667 |
| PA ~ Slope.l+Depth.c | 3 | -283.4283 | 527.8064 | 543.6773 | 20.2117 | 0.0000 | 9.1521 | 0.0000 | 36.8804 | 0.2974 | 0.9061 | 0.4669 | 0.8200 | 0.8667 |
| PA ~ Slope.l+JanSSTvar.c | 3 | -342.3985 | 651.7846 | 667.6555 | 131.7768 | 0.0000 | 120.6196 | 0.0000 | 23.7477 | 0.1662 | 0.8494 | 0.3533 | 0.7526 | 0.8074 |
| PA ~ Slope.l | 2 | -345.5559 | 658.1971 | 668.7805 | 136.8316 | 0.0000 | 120.6755 | 0.0000 | 23.0445 | 0.1621 | 0.8451 | 0.3483 | 0.7444 | 0.7852 |
| PA ~ Depth.c | 2 | -360.1929 | 696.0172 | 706.8605 | 168.8261 | 0.0000 | 152.4504 | 0.0000 | 19.7849 | 0.1376 | 0.8221 | 0.3257 | 0.6859 | 0.8000 |
| PA ~ JanSSTvar.c+JanSSTm.l+JanSSTm.l^2 | 4 | -378.5760 | 738.4619 | 759.6176 | 209.2229 | 0.0000 | 203.0834 | 0.0000 | 15.6909 | 0.0982 | 0.7945 | 0.2707 | 0.6919 | 0.7259 |
| MaxEnt (0.05) | 5 | -740.9223 | 1376.6198 | 1407.5116 | 0.0000 | 0.0000 | 0.0000 | 0.0053 | 23.2091 | 0.2625 | 0.9141 | 0.4847 | 0.7385 | 0.8741 |
| PA ~ Slope.l+JanSSTm+Depth+Slope.l^2+JanSSTm^2+Depth^2 | 6 | -737.9638 | 1374.3864 | 1414.6660 | 0.0000 | 0.1113 | 0.0000 | 0.0007 | 23.5157 | 0.2615 | 0.9139 | 0.4865 | 0.7474 | 0.8593 |
| PA ~ Slope.l+JanSSTvar.c+Slope.l^2+JanSSTvar.c^2 | 4 | -837.7823 | 1625.4989 | 1644.2664 | 172.0594 | 0.0000 | 156.9250 | 0.0000 | 13.1703 | -0.1734 | 0.8478 | 0.3563 | 0.4370 | 0.9481 |
| PA ~ Slope.l+Depth+Depth^2 | 4 | -842.2029 | 1636.3092 | 1646.2314 | 176.2320 | 0.0000 | 155.4890 | 0.0000 | 12.7121 | -0.2057 | 0.8439 | 0.3484 | 0.4281 | 0.9552 |
| PA ~ Depth+Depth^2 | 2 | -863.0877 | 1692.6760 | 1703.1619 | 223.1310 | 0.0000 | 199.9977 | 0.0000 | 10.5476 | -0.3832 | 0.8210 | 0.3271 | 0.3785 | 0.9704 |
| PA ~ JanSSTvar.c+JanSSTm+JanSSTvar.c^2+JanSSTm^2 | 4 | -885.0697 | 1742.1623 | 1763.0351 | 275.6037 | 0.0000 | 261.3210 | 0.0000 | 8.2693 | -0.7350 | 0.7959 | 0.2681 | 0.0696 | 1.0000 |
| BRT (0.05) | 7 | -956.6249 | - | - | - | - | - | - | 39.3689 | 0.2770 | 0.9010 | 0.4629 | 0.8466 | 0.7556 |
| PA ~ Slope.l+JanSSTm.l+Depth.c+JanSSTm.l^2 | 5 | -506.3007 | - | - | - | - | - | - | 39.4005 | 0.2749 | 0.9007 | 0.4609 | 0.8459 | 0.7556 |
| PA ~ Slope.l+Depth.c | 5 | -521.1883 | - | - | - | - | - | - | 37.5516 | 0.2681 | 0.8902 | 0.4586 | 0.8385 | 0.7556 |
| PA ~ Slope.l+JanSSTvar.c | 5 | -626.2371 | - | - | - | - | - | - | 26.5830 | 0.1533 | 0.8105 | 0.3434 | 0.7555 | 0.7407 |
| PA ~ Slope.l+dummy | 5 | -640.3154 | - | - | - | - | - | - | 25.0535 | 0.1483 | 0.8018 | 0.3353 | 0.7429 | 0.7556 |
| PA ~ Depth+dummy | 5 | -679.6601 | - | - | - | - | - | - | 21.0525 | 0.1257 | 0.7858 | 0.3170 | 0.6429 | 0.7481 |
| PA ~ JanSSTvar.c+JanSSTm.l+JanSSTm.l^2 | - | -964.2314 | - | - | - | - | - | - | 19.7063 | 0.0784 | 0.7609 | 0.2338 | 0.6570 | 0.7481 |
Table 3 95th percentiles of model metrics (Slope 1: log transformed bottom slope; JanSSTm 1: log-transformed January SST mean; JanSSTvar c: scaled interannual variance of January SST; Depth c: scaled bathymetric depth; JanSSTm: January SST mean; Depth: bathymetric depth; LogL: log-likelihood; AICc: a sample size corrected AIC; dAICc: the difference between AIC value of a candidate model and the smallest AIC value in the set of candidate models; wAICc: AICc weight; BIC: Bayesian Information Criterion; dBIC: the difference between BIC value of a candidate model and the smallest BIC value in the set of candidate models; wBIC: BIC weight; BIC differences; pcdev: percentage of explained deviance; cvweight: CV weight; r2n_pred: cross-validated R²)
