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RESEARCH ARTICLE

Perceptions of healthcare quality in Ghana: Does health insurance status matter?

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Abstract

This study’s objective is to provide an alternative explanation for the low enrolment in health insurance in Ghana by analysing differences in perceptions between the insured and uninsured of the non-technical quality of healthcare. It further explores the association between insurance status and perception of healthcare quality to ascertain whether insurance status matters in the perception of healthcare quality. Data from a survey of 1,903 households living in the catchment area of 64 health centres were used for the analysis. Two sample independent t-tests were employed to compare the average perceptions of the insured and uninsured on seven indicators of non-technical quality of healthcare. A generalised ordered logit regression, controlling for socio-economic characteristics and clustering at the health facility level, tested the association between insurance status and perceived quality of healthcare. The perceptions of the insured were found to be significantly more negative than the uninsured and those of the previously insured were significantly more negative than the never insured. Being insured was associated with a significantly lower perception of healthcare quality. Thus, once people are insured, they tend to perceive the quality of healthcare they receive as poor compared to those without insurance. This study demonstrated that health insurance status matters in the perceptions of healthcare quality. The findings also imply that perceptions of healthcare quality may be shaped by individual experiences at the health facilities, where the insured and uninsured may be treated differently. Health insurance then becomes less attractive due to the poor perception of the healthcare quality provided to individuals with insurance, resulting in low demand for health insurance in Ghana. Policy makers in Ghana should consider redesigning, reorganizing, and reengineering the National Healthcare Insurance Scheme to ensure the provision of better quality healthcare for both the insured and uninsured.
Introduction

Background

Healthcare financing in Sub-Saharan Africa (SSA) is a grim scenario. SSA comprises 12% of the world’s population and 22% of the total global disease burden [1], yet accounts for only 1% of the world’s health expenditure and 2% of the global workforce in healthcare. With low per capita income, limited domestic revenue mobilisation, and ineffective health systems, countries in SSA are ill-prepared to effectively address health financing problems [2]. Governments in most SSA countries are not the sole financiers of healthcare. In these cases, more than half of the health expenditure is financed through out-of-pocket payments, which places financial burdens on households and serves as a barrier to healthcare access [3, 4].

Countries in SSA, such as South Africa, Gabon, Mali, Senegal, Uganda, Tanzania, Nigeria, and Ghana have taken steps towards universal coverage by adopting risk pooling systems to provide financial protection, particularly to the poor and vulnerable in their societies [5–9]. These risk-pooling systems are either small, community-based health insurance (CBHI) schemes or social health insurance (SHI) schemes. The CBHI schemes are often voluntary and limited in geographic scope to a few rural communities with small numbers of enrollees as well as limited benefit coverage. The SHI schemes, on the other hand, have a wide scope, covering all regions in a country with subsidised premiums, government funding, and exemption policies. In some SHI schemes, enrolment is mandatory by law but voluntary in practice, as in Ghana. The benefit package covers quite comprehensive outpatient and inpatient healthcare services and medications [10, 11]. Surprisingly, enrolment in most of these SHI schemes remains low even though premiums are highly subsidised. In Ghana, as of December 2013, active membership in its National Health Insurance Scheme (NHIS) was 38% of the population [12]. People’s perception about the NHIS and healthcare quality has been identified as one of the factors that informs their decision to enrol or drop out of the scheme [13,14].

Prior economic literature has stipulated that the demand for health insurance is dependent on the quality of healthcare and assumes that quality is a constant, independent of health insurance status [15, 16]. The evidence on the relationship between health insurance enrolment and perceived quality of healthcare is very limited. Jehu-Appiah et al [14] found that the perceptions related to providers, schemes and community attributes play an important role at varying extents in household decision to voluntarily enrol and remain enrolled in insurance schemes. A systematic review by Spaan et al. [17] of the impact of health insurance in Africa and Asia, concluded that there is a weakly positive effect of SHI and CBHI on quality of healthcare and that the effect of health insurance on quality of healthcare is woefully under researched. However, recent qualitative evidence suggests that perceived quality of healthcare is not the same across insured and uninsured patients [18, 19]. A recent study by Robyn et al. [20], has also indicated that perceived quality of healthcare may be dependent on health insurance status, although they used a small sample of 398 patients to assess the effect of insurance status on technical quality and overall patient satisfaction.

This study builds on the work of Robyn et al. [20] to offer an explanation for the low demand for health insurance enrolment in Ghana, examining a larger sample and focusing on non-technical quality of care indicators (‘soft’ dimensions of quality), such as information provision, complaint lodging, and waiting times, instead of medical technical quality. The focus on non-technical quality of care stems from three main reasons. First, the NHIS accreditation process pays attention to medical technical quality and not non-technical quality. Second, the normative perspective of quality suggests that clients’ perceptions of quality are inherently meaningful and should be the primary focus of attention for quality assessments within healthcare systems. Finally, clients’ perceptions affect outcomes such as a health plan or health
provider choice, adherence to medical advice, complaints, and grievances; as well as health outcomes, which are powerful drivers important to stakeholders in the health sector [21–23].

We argue that enrolment levels in Ghana are low because there is a perception that service quality is lower for people with health insurance compared to those who make out-of-pocket payments. This study assesses whether the perceptions on healthcare quality depend on health insurance status by analysing cross-sectional household survey data of a random sample representative of 16 districts in two regions in Ghana. The survey included 1,903 household heads. Two-sample independent t-tests were used to compare the average perceptions of quality per health facility between insured and uninsured respondents. A generalised ordered logit regression, controlling for socio-economic characteristics and clustering at health facility level, was done to test the association between insurance status and perceived quality of healthcare.

The health system and NHIS financing in Ghana

The responsibility of the Ministry of Health in Ghana is to provide policy guidance, regulation, and strategic direction to service providers, regulatory bodies, and the National Health Insurance Authority (NHIA) [24, 25]. The service providers include Ghana Health Service (responsible for service delivery in all public health facilities), non-governmental organizations (e.g., Christian Health Association of Ghana), and private providers. There are however more private health facilities in the urban areas as compared to the rural areas and vice versa. The quality of healthcare in private health facilities is generally purported to be higher than public ones, whiles the public facilities are often regarded as providing equitable and evidence based healthcare. For this reason, the rich who can afford out-of-pocket payments and private health insurance often patronize private facilities while the poor, mostly in rural areas patronize public facilities [26]. The NHIA and other regulatory bodies (e.g., the Medical & Dental Council) assess and regulate standards of service quality. The NHIA accreditation process assesses the technical quality of the health facilities before their accreditation to provide healthcare services. These technical assessments focus on medical technical quality with no attention to non-technical quality indicators such as interpersonal relations, queuing systems, complaints, and waiting times [25].

The NHIS was established in 2004 to ensure financial access to equitable and acceptable quality of essential healthcare benefits for all residents in Ghana. The scheme is financed mainly through the National Health Insurance Fund (NHIF). The NHIF monies come from 2.5% of the 17.5% value-added tax (VAT), 2.5% of the 17.5% Social Security and National Insurance Trust (SSNIT), contributions from formal sector employees, accruals to the fund from investments of the NHIA Council, and member contributions from premium payments. The annual premium is set at the NHIS district office level and is dependent on the poverty and economic activity levels of the districts with approval from the NHIS council. The average annual premium is about GHC25 ($13.14 at an exchange rate of GHC1 = $1.87 as at December, 2012). The Government of Ghana also allocates funds to the NHIF through parliament and other donor funds [27–30, 24]. The NHIS has a broad benefit package, covering 95% of the burden of diseases in Ghana. The NHIS covers out-patient services (consultation including reviews for general and specialist outpatient consultations); in-patient services (general and specialist in-patient care); maternity care (including antenatal care, normal or assisted deliveries); eye care (refraction, visual fields, A-scan, keratometry) and oral health (pain relief, drainage, tooth extraction, dental restoration, fillings and temporary dressing). It provides premium exemptions for the elderly (70 years and above), SSNIT pensioners, children below 18 years, indigents, pregnant women, and Livelihood Empowerment Against Poverty (LEAP) beneficiaries [11]. Nevertheless, the NHIS claims account for only 16% of total healthcare expenditure.
and 30% of public health expenditure, while out-of-pocket spending accounts for 37% of total health spending. This exceeds the World Health Organisation (WHO) suggested threshold of 15–20% for out-of-pocket expenditure for adequate financial protection [25].

At the inception of the NHIS, the payment of healthcare providers was by itemized fee-for-service. However, in 2008, the NHIA reformed the provider payment mechanism and introduced the Ghana Diagnostic Related Groupings (G-DRGs) for services and fee-for-service for medicines at all levels of service delivery. The NHIS financing arrangement is such that NHIS district offices sign purchasing contracts with accredited healthcare providers. These contracts stipulate the prescription of generic drugs for the treatment of specific medical conditions and healthcare providers are to submit claims for services rendered to insured clients to the NHIS district offices for reimbursement.

Conceptual framework

People’s perception of healthcare quality is the outcome of an evaluation process where expectations are compared to actual experiences and realities of care received [31, 32]. Individuals enrol in health insurance because they expect financial protection from excessive out-of-pocket payments for the provision of quality healthcare. The consensus among researchers is that positive perceptions of healthcare quality lead to increased client satisfaction, and the acknowledgment of value and trust in the healthcare provider [32, 16, 15, 33, 34]. This ultimately affects individuals’ demand for healthcare and hence, health insurance.

The utilisation of healthcare is found to be sensitive to quality of care such that households limit their demand when services are of poor quality [35–37] and others bypass low quality health facilities in search of high quality ones [38–41]. The generally accepted interpretation that perception of high quality healthcare influences healthcare utilisation and insurance enrolment assumes that perceptions are independent of insurance status as shown in the “likely pathways” in Fig 1 and ignores a potential dependent effect in the other direction; that is, once insured, people may perceive their healthcare quality as poor depending on the service delivery processes they go through. This may be due to dissatisfaction with the service delivery processes as well as poor technical quality of care provided to the insured. The insured are reported to indicate lower perceived quality of care than the uninsured in recent studies [18, 19]. This paper argues that perception on healthcare quality may be dependent on health insurance status as shown in the likely pathways in Fig 2 below. Hence, ‘enrolment status (insured or uninsured) can influence the direction (low or high) of quality perception’.

If the insured perceive the quality of the healthcare they receive to be low, how does that influence their health insurance renewal decisions. This study contributes the sparse evidence in the literature on the dependence of perceived quality of healthcare on health insurance status.

Methods

Study setting

This study is part of a larger research project: the Client-Oriented Health Insurance System in Ghana (COHEiSION) Project. The aim of the research is to identify the main perceived barriers to (re)-enrol in the NHIS, and to design, implement, and evaluate the most effective intervention that can address them. The research was conducted in the Greater Accra and Western regions of Ghana. Ideally, a mix of one northern and one southern region would have provided a much better balance in terms of health insurance enrolment coverage and socio-economic setting. However due to resource and time constraints these two southern regions were selected. These regions are situated in the Eastern and Western coast of Ghana, respectively,
with a contrasting difference in urban and rural populations. The Greater Accra Region with Accra, the capital city of Ghana, has a largely urban population of about four million, accounting for 16.3% of the national population. The Western region has a predominantly rural population of about two million, representing 9.6% of the national population [42]. These two regions were purposively selected to provide the rural/urban balance which was of interest to the Project.

Sampling procedure and data collection

This study analyses data that are representative of 64 communities in 16 districts from the baseline survey of the research project based on a three-stage stratified sampling procedure. At the first stage, 16 districts were purposively selected from the two regions, eight in each region. The second stage involved the purposive selection of 64 primary healthcare facilities (government...
health centres and private clinics), four in each of the 16 districts. At the third stage, enumeration area (EA) maps were used to list all residential buildings within a 10 km radius of the catchment area of each selected health facility. Simple random sampling was then used to sample 30 households from the catchment area of each of the 64 health facilities to obtain a total sample of 1,920 households.

A baseline household survey was conducted from March to April 2012. The survey used a paper based structured questionnaire, which was translated into the two local languages (Ga and Fante) of the project regions. The questionnaire was administered to household heads by experienced and trained interviewers. The survey collected information on socio-demographics, employment status, consumption expenditure patterns, health status, healthcare utilisation behaviour, dwelling characteristics, NHIS enrolment status, and perceptions of non-technical quality of healthcare.

Measurement of variables

The dependent variables of interest were perceptions on healthcare quality. These were measured by asking household heads to rate on a five-point Likert scale (1 = very dissatisfied; 2 = dissatisfied; 3 = neutral; 4 = satisfied; 5 = very satisfied or 1 = strongly disagree; 2 = disagree; 3 = neutral; 4 = agree; 5 = strongly agree) their satisfaction or agreement with seven perception questions about the non-technical quality of NHIS accredited primary healthcare facilities. For each respondent, an overall healthcare quality index was calculated as an average score of the seven perceptions of quality indicators. The distribution of the average scores was fit into five quintiles to create a five-level ordinal scale for the overall healthcare quality index. The seven quality indicators were (1) satisfaction with service provision; (2) satisfaction with process of lodging a complaint; (3) satisfaction with information provision; (4) satisfaction with waiting time; (5) agreement with the statement that all prescribed drugs were received from the facility; (6) agreement with the statement that health personnel treated patients with insurance cards equally to those without cards and; (7) agreement with the statement that there was a fair queuing system at the health facility.

The key independent variable of interest was insurance status. This variable was measured by asking respondents to indicate Yes = 1, if they were currently insured in any health insurance scheme, and No = 2, if not. The currently insured are respondents who had valid health insurance cards at the time of the survey and the currently uninsured are those who did not. The currently uninsured were further required to indicate Yes = 1, if they were previously insured in any health insurance scheme and No = 2, if they have never been insured in any health insurance scheme. The previously insured are respondents who had insurance cards but was not valid for them to access healthcare with at the time of the survey because they did not renew their membership in the year 2012, whiles the never insured are those who have never enrolled in any health insurance before. Respondents’ socio-demographic characteristics were also captured by asking them to state their age, sex, marital status, religion, level of complete education, size of household, rural or urban nature of locality, employment status, self-assessed health status, and weekly consumption expenditure (S1 Appendix).

Ethical considerations

Ethical clearance for the research was obtained from the Ghana Health Service (GHS) Ethical Review Committee (ERC) with ethical clearance number: GHS-ERC: 08/5/11. The ethical clearance approved for informed consent to be obtained from the household heads. Literate respondents provided written informed consent, while illiterate respondents thumb-printed the informed consent form read to them in their local language before participating in the study. Data was also analysed anonymously.
Statistical analysis

All statistical analysis was performed with Stata software version 12.1. First, the averages of the respondents’ demographic and socio-economic statistics were calculated. The differences in these descriptive statistics between the insured and uninsured, and also between the previously insured and never insured, was done using a two-sample independent t-test.

Two-sample independent t-tests were then used to test the hypothesis that there is no difference in the perception on healthcare quality between the insured and uninsured; or between the previously insured and never insured. Given the ordinal five-level categories of the dependent variables, ordinal logistic regression (proportional odds/parallel-lines model) would have been a simple straightforward, intuitive, and easy to interpret method to estimate the effect of insurance status on perception of non-technical quality of healthcare. However, the main problem with the parallel-lines model is that the proportional odds assumption is often violated. To ascertain this, we used the *brant* command in Stata, which does both a global test of whether the independent variables violate the proportional odds assumptions as well as a test of whether any of the independent variables separately violate the assumptions. We obtained significant test statistics for both the model estimation as well as some of the independent variables, an indication that the proportional odds assumptions of the model had been violated. To address this problem, we employed the less restrictive generalised ordered logit model (partial proportional odds model), which relaxes the proportional odds assumption and allows the effect of the independent variables to vary with the point at which the categories of the dependent variable are dichotomised. The standard formula for the predicted probability in the partial proportional odds model is denoted as:

\[
P(Y_i < j) = \frac{\exp(z_i \beta_1 + X_i \beta_2 + X_i \beta_3)}{1 + \{\exp(z_i \beta_1 + X_i \beta_2 + X_i \beta_3)\}} \quad j = 1, 2, \ldots, M - 1
\]

where \(Y\) is the ordinal dependent variable and \(M\) is the number of categories of the ordinal dependent variable [43]. The regression estimation was done for the seven indicators of healthcare quality and the overall quality index as well as controlling for key demographic and socio-economic variables [20].

The key independent variable (explanatory variable) of interest in the regression estimation is health insurance status. This variable was coded with a dichotomous response, equal to 1 if the respondent is currently insured, and 0 if currently uninsured, or equal to 1 if the respondent was never insured, and 0 if previously insured. The demographic variables used in the regression estimation as controls were age (continuous), gender (male = 0, female = 1), married or not (not married = 0, married = 1), religion, (Other religion = 0, Christian = 1), residential zone (rural = 0, urban = 1), number of people in the household (continuous), and number of health facility visits (continuous). Other socio-economic variables included as controls were level of completed education (below primary level = 0, primary education and above = 1) and employment status (unemployed = 0, employed = 1).

Results

Descriptive characteristics and health insurance enrolment

Of the 1,920 households sampled, the baseline survey interviewed 1,903 household heads (948 in the Western region and 955 in the Greater Accra region), representing a 99.1% interview consent rate; 17 household heads declined to participate in the survey.

The descriptive characteristics of respondents and the differences in these characteristics between the currently insured and currently uninsured and between the previously insured
and never insured are presented in Table 1. Approximately 40% of the respondents were enrolled in health insurance at the time of the survey and of the uninsured, approximately 70% had never been insured in any health insurance scheme.

The average age of respondents was approximately 45 years old with 36.4% being female. Slightly more than half (52.9%) were married with the majority of them (88.6%) being Christian. About half (50.2%) of the respondents lived in rural communities with an average household size of approximately four people. The currently insured were older (48 years), with slightly more female (37%) and more married (56.1%) respondents, equally likely to live in rural (50.2%) or urban communities, and more likely to have a household close to the average size (3.9) compared to the currently uninsured. These differences in the demographic characteristics between the currently insured and currently uninsured were statistically significant at the 95% confidence level except for the percentage of females and the percentage living in rural communities. Similarly, the previously insured were significantly more likely to have higher averages and proportions in these demographic characteristics than the never insured, except for the married percentage and percentage living in rural communities (See Panel A of Table 1).
Most (73.2%) of the respondents completed a primary or above level of education with the majority (86.3%) gainfully employed. The average annual household consumption expenditure was GHC3,685.40 ($2,538.61) in 2012. Although the currently insured (76.6%) were significantly more likely to have completed a primary or above level of education, the currently uninsured (88.3%) were more likely to be gainfully employed and also more likely to spend more (GHC3,847.30 or $2,650.13) on household consumption. Except for the annual household consumption expenditure, these differences between the currently insured and currently uninsured were statistically significant at the 95% confidence level. The previously insured, however, were less likely to have higher averages and proportions in these socio-economic characteristics than the never insured, except for the percentage with primary or above level of education where both the previously insured and never insured were equal (see Panel B of Table 1).

A high proportion (84.6%) of respondents indicated that they were in good health; in the six months prior to the household survey, the average number of illnesses was 1.1 and the average number of health facility visits was 0.8. The currently insured (78.7%) were significantly less likely to report being in good health; they were significantly more likely (1.3 times) to report being ill and more likely (1.1 times) to visit a health facility. The previously insured (87.2%) were, however, less likely to report being in good health but equally likely (0.9 times) to report being ill and more likely (0.6 times) to visit a health facility compared to the never insured (see Panel C of Table 1).

**Perception of non-technical quality of healthcare**

Generally, the average perception of healthcare quality among the respondents was relatively low, with the highest average perception approximately three on the scale (equivalent to neutral on the five-point Likert scale) for the ‘process of lodging complaints at the health facility’.

The currently uninsured perceived the quality of healthcare to be better than the currently insured on all seven quality indicators. The average perception of the currently uninsured (2.09, p-value = 0.000; 2.81, p-value = 0.018; 2.21, p-value = 0.000; 2.36, p-value = 0.000; 2.48; 2.59, p-value = 0.000; 2.08, p-value = 0.000, and 2.33, p-value = 0.000) was significantly higher than that of the currently insured on: services provided; process of lodging complaints; information provision; waiting time; availability of prescribed drugs; equal treatment of insured and uninsured patients; fair queuing system; and, overall average perception index, respectively.

Similarly, when the currently uninsured were disaggregated into the previously insured and never insured, the perceptions of the never insured (2.15, p-value = 0.000; 2.85, p-value = 0.351; 2.26, p-value = 0.005; 2.64, p-value = 0.033; 2.08, p-value = 0.824 and 2.36, p-value = 0.023) were found to be higher than the previously insured on: services provided; process of lodging complaints; information provision; waiting time; equal treatment of insured and uninsured patients; fair queuing system; and, overall average perception index, respectively.

Table 2 presents the differences in the average perceptions on healthcare quality between the currently insured and currently uninsured, and between the previously insured and never insured.

However, for the perception on availability of prescribed drugs at the health facility (2.5, p-value = 0.721), the perceptions of the previously insured were higher than that of the never insured, but the difference was statistically insignificant at the 95% confidence level.

**Effect of current insurance status on perceived quality of healthcare**

Table 3 presents the results of the generalised logit regression estimation of the effect of current health insurance status on the perception of healthcare quality. Only the odds ratio of the independent variable of interest (insurance status) and the constant are reported for the various
ratings of quality perceptions. The entire results is however presented in S2 Appendix. The currently insured were significantly less likely than the currently uninsured to choose strongly disagree or very dissatisfied for the average perceptions (OR = -0.264) on healthcare quality. Similarly, the currently insured were significantly less likely to report strongly disagree or very dissatisfied (OR = -0.467; -0.294; -0.485; -0.451; -0.346 & -0.285) with the seven indicators (perceptions on service provision, complaint lodging, information provision, waiting time, prescribed drugs, equal treatment of insured and uninsured, and queuing system). This can be seen from the negative currently insured coefficients for the strongly disagree and very dissatisfied panel in Table 3.

The currently insured were also less likely than the currently uninsured to choose disagree or dissatisfied for their average perception (OR = -0.705) on healthcare quality (OR = -0.731, -0.294, -0.755, -0.531, -0.451, -0.445 & -0.518) for the seven indicators (service provision, complaint lodging, information provision, waiting time, prescribed drugs, equal treatment, and queuing system) of quality of healthcare (see the negative currently insured coefficients in the disagree and dissatisfied panel of Table 3).

Although the currently insured were less likely than the currently uninsured to choose neutral for their average perception or for the seven indicators of healthcare quality, it was only the results of their perception on complaint lodging (OR = -0.294) that was statistically significant (see the negative insured coefficients in the neutral panel of Table 3). The currently insured were, however, found to be less likely than the currently uninsured to choose agree or satisfied with the statements about prescribed drugs (OR = -0.069), equal treatment (OR = -0.262), and queuing system (OR = -0.118). These findings were statistically insignificant. However, In terms of the magnitude, the effect of insurance status was less extreme at the lower levels of the perceived healthcare quality categories. Thus, as one moves from disagree

Table 2. Differences between the insured and uninsured in the average perception of quality of 64 health facilities.

|                          | Total Obs. | Mean | Std. Error | Currently Insured | Currently Uninsured | Diff. | p-value (t-test) | Previously Insured | Never Insured | Diff. | p-value (t-test) |
|--------------------------|------------|------|------------|-------------------|---------------------|-------|----------------|-------------------|----------------|-------|----------------|
| How satisfied are you with the services provided by the health facility? | 1869       | 1.99 | 0.02       | 1.84              | 2.09                | 0.24  | 0.000          | 1.94              | 2.15           | -0.21 | 0.000          |
| How satisfied are you with the process of lodging complaint at the facility? | 1866       | 2.78 | 0.02       | 2.73              | 2.81                | 0.08  | 0.018          | 2.78              | 2.85           | -0.04 | 0.351          |
| How satisfied are you with the information provided by the health facility? | 1874       | 2.11 | 0.02       | 1.95              | 2.21                | 0.26  | 0.000          | 2.09              | 2.26           | -0.16 | 0.005          |
| How satisfied are you with the waiting time at the facility? | 1789       | 2.27 | 0.02       | 2.14              | 2.36                | 0.23  | 0.000          | 2.26              | 2.40           | -0.14 | 0.042          |
| I received all prescribed drugs from the facility | 1879       | 2.41 | 0.03       | 2.31              | 2.48                | 0.17  | 0.006          | 2.50              | 2.47           | 0.03  | 0.721          |
| Health personnel treats patients with insurance cards in an equal way as patients without cards | 1884       | 2.50 | 0.03       | 2.24              | 2.59                | 0.22  | 0.000          | 2.46              | 2.64           | -0.18 | 0.033          |
| There is a fair queuing system at the health facility. | 1884       | 2.01 | 0.02       | 1.90              | 2.08                | 0.18  | 0.000          | 2.07              | 2.08           | -0.01 | 0.824          |
| Overall Average Perception | 1886       | 2.26 | 0.02       | 2.14              | 2.33                | 0.19  | 0.000          | 2.25              | 2.36           | -0.11 | 0.023          |

Source: COHEIISON Project baseline survey (March 2012), N = 1,903 household heads. Note: Standard errors are robust and corrected for clustering at the health facility level.

*: p<0.10
**: p<0.05
***: p<0.01

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or dissatisfied to strongly disagree or very dissatisfied, the likelihood of the currently insured being in that category increases in magnitude.

We also found that married respondents were less likely to choose strongly disagree, very dissatisfied, disagree or dissatisfied, whereas those who made health facility visits and those with primary level education or above were significantly more likely to report strongly disagree, very dissatisfied, disagree, and dissatisfied with quality of healthcare.

**Effect of previous insurance status on perceived quality of healthcare**

Table 4 presents the generalised ordered logit regression estimation of the effect of never insured status on the perception on healthcare quality. Again only the odds ratio of the independent variable of interest and the constant are reported for the various ratings of quality perceptions and the entire results presented in S3 Appendix. The results show that although the never insured are more likely to choose strongly disagree or very dissatisfied and disagree or dissatisfied for their average perception (OR = 0.245) healthcare quality, this finding was statistically insignificant. The never insured were, however, significantly more likely than the previously insured to choose strongly disagree or very dissatisfied for their perception of service provision (OR = 0.330), information provision (OR = 0.294), and equal treatment (OR = 0.418) (see the positive never insured coefficients in the strongly agree or strongly dissatisfied panel in Table 4).

The never insured were also more likely than the previously insured to choose disagree or dissatisfied for perceptions on service provision (OR = 0.709), information provision (OR = 0.556),

### Table 3. Effect of currently insured status on perceived quality of healthcare.

|                      | Average Perception | Service Provision | Complaint Lodging | Information Provision | Waiting Time | Prescribed Drugs | Equal Treatment | Queuing System |
|----------------------|--------------------|-------------------|-------------------|----------------------|--------------|------------------|----------------|----------------|
| **Strongly Disagree or Very Dissatisfied** Insured | -0.264*             | -0.467***          | -0.294**          | -0.485***           | -0.451***    | -0.150           | -0.346***       | -0.285**       |
|                      | (0.154)            | (0.122)            | (0.117)           | (0.124)             | (0.128)      | (0.107)          | (0.121)         | (0.112)         |
| _cons                | 1.950***           | 0.708***           | 2.569***          | 1.341***            | 1.346***     | 1.376***         | 1.047***        | 0.823***        |
|                      | (0.262)            | (0.262)            | (0.408)           | (0.275)             | (0.269)      | (0.292)          | (0.274)         | (0.258)         |
| **Disagree or Dissatisfied** Insured | -0.705***          | -0.731***          | -0.294**          | -0.755***           | -0.531***    | -0.451***        | -0.445***       | -0.518***       |
|                      | (0.131)            | (0.145)            | (0.117)           | (0.140)             | (0.112)      | (0.108)          | (0.105)         | (0.124)         |
| _cons                | -0.103             | -0.800***          | 1.751***          | -0.408              | 0.132        | 0.395            | 0.116           | -0.529**        |
|                      | (0.272)            | (0.252)            | (0.396)           | (0.258)             | (0.262)      | (0.277)          | (0.272)         | (0.269)         |
| **Neutral** Insured  | -0.088             | -0.136             | -0.294**          | -0.248              | -0.173       | -0.100           | -0.172          | 0.024           |
|                      | (0.190)            | (0.186)            | (0.117)           | (0.176)             | (0.145)      | (0.111)          | (0.124)         | (0.175)         |
| _cons                | -2.495***          | -2.299***          | -2.488***         | -2.289***           | -1.683***    | -0.697**         | -1.376***       | -1.960***       |
|                      | (0.315)            | (0.298)            | (0.364)           | (0.327)             | (0.268)      | (0.308)          | (0.294)         | (0.302)         |
| **Agree or Satisfied** Insured | -0.069             | -0.262             | -0.118            |                     |             |                  |                |                |
|                      | (0.166)            | (0.161)            | (0.282)           |                     |             |                  |                |                |
| _cons                | -1.747***          | -2.147***          | -2.547***         |                     |             |                  |                |                |
|                      | (0.309)            | (0.320)            | (0.329)           |                     |             |                  |                |                |
| **No. of Obs.**      | 1785               | 1768               | 1765              | 1774                | 1694         | 1778             | 1783            | 1783            |

Source: COHEIISON Project baseline survey (March 2012), N = 1,903 household heads. Note: Standard errors in parenthesis are robust and corrected for clustering at the health facility level.

*: p<0.10  
**: p<0.05  
***: p<0.01.

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equal treatment (OR = 0.451), and queuing system (OR = 0.319) (see the positive never insured coefficients in the disagree or dissatisfied panel in Table 4). In terms of magnitude, the never insured effect was more extreme at lower levels of the perceived healthcare quality ratings. This means that as one moves from disagree or dissatisfied to strongly disagree or very dissatisfied, the likelihood of the never insured being in that category reduces in magnitude.

The never insured were, however, found to be more likely than the previously insured to choose neutral for the average perception, service provision, complaint lodging, waiting time, and equal treatment, but less likely than the previously insured to choose neutral for their perceptions of information provision and prescribed drugs. However, these findings were statistically insignificant (see the positive and negative never insured coefficients in the neutral panel in Table 4). The never insured were significantly less likely to choose agreed or satisfied for only their perception of prescribed drugs (OR = 0.633), as shown in the negative never insured coefficient of the agree or satisfied panel in Table 4.

**Discussion**

**Perception of non-technical quality of healthcare**

We found that within the NHIS in Ghana, there were significant differences between the currently insured and currently uninsured; and between the previously insured and never insured
Health insurance status and perceptions of healthcare quality

Effect of insurance status on perception of healthcare quality

Health insurance enrolment had a significant negative effect on the perception on healthcare quality. The regression estimation indicates that being currently insured is associated with a significantly lower perception on healthcare quality. When the currently uninsured were disaggregated into never insured and previously insured, being never insured was associated with a significantly higher perception of most of the healthcare quality indicators. This means that once people are insured, they tend to perceive the quality of healthcare to be poor compared to those uninsured, whereas those who have never been insured before tend to perceive the quality of healthcare as better than those who once had health insurance and dropped it. This finding is, however, in contrast with that of Robyn et al. [20] who found that both the insured and uninsured in the Nouna District of Burkina Faso gave high ratings for quality of healthcare delivery.

A plausible explanation for this could be that under the NHIS, insured clients actually receive lower levels of actual quality due to the differences in the healthcare delivery process for the insured and uninsured at the health facilities. The service delivery processes for NHIS accredited health providers are such that they have to process claims forms for consultation, diagnosis, and drugs for insured clients. However, for the uninsured, providers only have to provide care and collect cash. Therefore, the uninsured are kept in a different queue, quickly seen at the out-patients’ department for consultation and at the pharmacy for drugs, whereas the insured spend long hours in queues to go through the NHIS process for receiving healthcare. The uninsured, therefore, do not experience the long waiting times that the insured experience. This coupled with the long delays, in excess of over six months, in NHIS reimbursement, creates a situation where healthcare providers have no incentive to invest in the provision of quality services for insured clients when they can provide services to uninsured clients and receive cash immediately. This translates into the uninsured being satisfied with the waiting times, and the insured feeling they spend long hours in queues at the health facilities, as found in this study, and also documented by Fenenga et al. [18].

The NHIS purchasing contracts with accredited providers specifies the prescription of generic drugs for insured clients. Thus, providers who dispense branded drugs are not reimbursed. This, in addition to the delays in the reimbursement of claims, serves as a further disincentive for providers to stock generic drugs. Insured clients are asked to buy generic drugs from outside the facility, whilst uninsured clients pay cash for branded drugs at prevailing
market prices at the health facilities. Insured clients interpret their prescriptions of generic drugs and branded drugs for the uninsured as discriminatory, as documented by Jehu-Appiah et al. [33] and Fenenga et al. [18]. These differences in the healthcare delivery process for the insured and uninsured might represent a lower level of actual quality to insured clients, which translates into their negative perceptions of healthcare quality as found in this study.

**Conclusion**

This study demonstrated that health insurance status matters in the perception of the non-technical quality of healthcare in Ghana. The findings suggest that people’s perceptions of healthcare quality may be shaped by their actual experiences at the health facilities, and these experiences differ depending on their insurance status. The insured who experience long delays and are served with generic drugs due to the different services delivery process they go through perceive the quality of care they receive to be low whiles the uninsured who pay cash for branded drugs and spend less time perceive the quality of the care they receive to be relatively higher. The implication is that enrolment levels in the NHIS are low because service quality is perceived to be lower for people with health insurance compared to those who make out-of-pocket payments. The insured who are dissatisfied with the quality of care may drop out because they think they are better off opting for out-of-pocket payments, while others may renew because they believe the financial protection of not having to make out-of-pocket payments over-rides the low quality of care. Therefore, it is important that policy makers consider redesigning, reorganising, and reengineering the NHIS, particularly the provider payment mechanism, to ensure the provision of quality healthcare services for all.

The authors acknowledge some limitations that might have affected the results. The sample for the study was drawn from two of the 10 regions in Ghana. The sample size might therefore not be representative of the Ghanaian population. The survey questionnaire was translated from English to the two local languages (Ga and Fante) of the study regions for the respondents in the rural communities to understand and thus could be affected by interviewer bias. The perception questions could lose their actual meaning in the process. The perceptions of respondents may differ significantly in the other regions of the country due to differences in the number of public and private health facilities in these other regions. The two groups (insured and uninsured) were significantly different in terms of age, marital status, religion, educational level, employment status, and health status, which were controlled for in the analysis. It is possible that these two groups also differ on other unobserved factors, which could not be controlled for in the analysis, and therefore, the observed effect of insurance status on perceived quality of healthcare could be the result of systematic differences between the two groups or other unobserved factors. These limitations notwithstanding, the findings from this study remain relevant to the NHIS, the Ghana Health Service, and other countries implementing social health insurance schemes for the provision of quality healthcare services to the insured and uninsured and the sustainability of such schemes.

**Supporting information**

S1 Appendix. Baseline survey questionnaire.

(OCR)

S2 Appendix. Effect of currently insured status on perceived quality of healthcare.

(OCR)

S3 Appendix. Association of never insured status on perceived quality of healthcare.

(OCR)
S4 Appendix. Perceived quality data.
(DTA)

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