Bayesian analysis of inflationary features in Planck and SDSS data

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We perform a Bayesian analysis to study possible features in the primordial inflationary power spectrum of scalar perturbations. In particular, we analyse the possibility of detecting the imprint of these primordial features in the anisotropy temperature power spectrum of the Cosmic Microwave Background (CMB) and also in the matter power spectrum $P(k)$. We use the most recent CMB data provided by the Planck Collaboration and $P(k)$ measurements from the eleventh data release of the Sloan Digital Sky Survey. We focus on a class of potentials whose features are localised at different intervals of angular scales, corresponding to multipoles in the ranges $10 < \ell < 60$ (Oscill-1) and $150 < \ell < 300$ (Oscill-2). Our results show that one of the step-potentials (Oscill-1) provides a better fit to the CMB data than does the featureless ΛCDM scenario, with a moderate Bayesian evidence in favor of the former. Adding the $P(k)$ data to the analysis weakens the evidence of the Oscill-1 potential relative to the standard model and strengthens the evidence of this latter scenario with respect to the Oscill-2 model.

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I. INTRODUCTION

The inflationary paradigm offers an elegant theoretical framework in which the emergence of the primordial curvature perturbations can be understood. In the simplest inflationary scenarios a primordial scalar fluctuation with nearly scale-invariant power spectrum is generated by a single minimally-coupled scalar field $\phi$ rolling down a smooth potential $V(\phi)$. This framework seems to agree with the most recent cosmic microwave background (CMB) data [1][2], which show a preference for plateau-like potentials over monomial potentials, as well as no compelling statistical evidence for a specific class of scenarios (we refer the reader to [7][11] for different points of view of the current observational status of inflation).

Recently, several works have analysed inflationary models that account for localised features in the primordial power spectrum, showing in some cases a better fit to the data with respect to a smooth power-law spectrum [12][22]. Features in the primordial power spectrum can be generated following departures from the slow-roll approximation, which can happen in more general inflationary scenarios with a symmetry-breaking phase transition. Examples are the inflationary models with a step in the primordial potential, whose oscillation in the power spectrum of curvature perturbations is localised around the scale that crosses the horizon at the time the phase transition occurs [23][24].

In principle, this oscillation could be seen in several observables, such as in the CMB maps and in the large-scale galaxy distribution, since the primordial scalar fluctuation is the seed for all the cosmic structures currently observed. It is therefore expected that data of the temperature anisotropy power spectrum as well as measurements of the matter power spectrum $P(k)$ from galaxy surveys may provide hints on the origin of these features. Originally, step-like inflationary potentials were postulated to study the “glitch” appearing at the low-$\ell$ region of the CMB anisotropy spectrum. The signatures of this class of models in the CMB temperature power spectrum and bispectrum [17][21][25][31] and in the tensor spectrum [29][33] have been studied in detail using the current data and have shown a consistent improvement of the $\Delta \chi^2$ values with respect to the featureless ΛCDM model.

In this work we proceed a step further in this kind of analysis and employ a Bayesian statistical analysis to verify the predictions of a class of inflationary step-like potential models and discuss their observational viability in a wide range of scales, considering not only their observational consequences on the present CMB anisotropy temperature maps but also on the matter power spectrum $P(k)$. We work with two data sets: the second CMB data release of the Planck Collaboration [2][34] and a combination of these CMB data with measurements of the matter power spectrum from the Baryon Oscillation Spectroscopic Survey (BOSS) CMASS Data Release-11 sample of the Sloan Digital Sky Survey (DR11-SDSS) experiment (hereafter “CMB+SDSS”) [37]. We perform a Bayesian analysis using both the Metropolis-Hastings algorithm implemented in CosmoMC [43] and the nested sampling algorithm of MultiNest [14][15]. We find that at least one of the scenarios studied is able to provide a better fit to the CMB and CMB + $P(k)$ data than does the standard ΛCDM model.

This paper is organised as follows. Sec. [1] reviews the class of inflationary models considered in this work. We also discuss observational data sets and priors used in the analysis as well as the Bayesian model selection method adopted. In Sec. [11] we discuss the results and present

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1 For plateau-like potentials $V(\phi) \to 0$ as $\phi \to \infty$ (see, e.g., [3][25]).
a brief comparison with previous analysis. We end the paper by summarising the main results in Sec. IV.

II. MODEL AND METHOD

Relaxing the slow-roll condition leaves traces on the primordial power spectrum. For example, if the wavelengths cross the horizon during the fast-roll phase, one must expect deviations from the usual power-law power spectrum. The brief violation of the slow-roll condition can be shaped, in a single-field model, by adding a local feature, such as a step, to an otherwise flat potential. Following the formalism presented in Refs. [23][41][42], we consider a model with a local feature added to a chaotic potential

\[ V(φ) = \frac{1}{2} m^2 φ^2 \left[ 1 + c \tanh \left( \frac{φ - b}{d} \right) \right]. \] (1)

The spectrum of primordial perturbations, resulting from the potential (1), is found to be essentially a power-law with superimposed oscillations. These are centred on a value that depends on the parameter \( b \), with amplitude set by \( c \) and damping given by \( d \). Asymptotically, the spectrum recovers the familiar \( k^{n_s-1} \) form, typical of slow-roll inflationary models.

In this work we adopt an analytical parametrisation for the scalar primordial spectrum resulting from Eq. (1), as studied in [25]:

\[ P_\ell(k) = \exp \left[ \ln P_0(k) + \frac{A_f}{3} \frac{k \eta_f}{\sinh \left( \frac{k \eta_f}{x_d} \right)} W'(k \eta_f) \right], \] (2)

where

\[ W'(x) = (-3 + \frac{9}{2x}) \cos 2x + (15 - \frac{9}{2x}) \sin 2x, \] (3)

\[ P_0(k) = A_s (\frac{k}{Mpc})^{n_s-1} \]

is the smooth spectrum with the standard power-law form, \( A_s \) is the kinetic energy perturbation of the step, \( \eta_f \) is the step crossing time in units of Mpc and \( x_d \) is the dimensionless damping scale. The oscillating window function \( W'(k \eta_f) \) is modulated by the decaying envelope \( k \eta_f / \sinh(k \eta_f / x_d) \) which is set by the details of the step. It is worth mentioning that the approximate approach of Eqs. (2) and (3) allows to significantly save computing time with respect to the full numerical solution of the evolution equations for the potential of Eq. (1), without a meaningful loss of accuracy [20].

We consider a “vanilla” model with the addition of features in the primordial spectrum, parametrised as in Eqs. (2) and (3). In our analysis, we vary the usual cosmological parameters, namely, the physical baryon density, \( Ω_b h^2 \), the physical cold dark matter density, \( Ω_c h^2 \), the ratio between the sound horizon and the angular diameter distance at decoupling, \( θ_s \), the optical depth, \( τ \), the primordial scalar amplitude, \( A_s \), the primordial spectral index, \( n_s \), and the additional \( A_f \), \( η_f \) and \( x_d \) step parameters. We also vary the nuisance foreground parameters [34] and consider purely adiabatic initial conditions. The sum of neutrino masses if fixed to 0.06 eV, and we limit the analysis to scalar perturbations with \( k_0 = 0.05 \) Mpc\(^{-1} \).

The posterior probabilities distributions of the parameters are generated using both the Metropolis-Hastings algorithm implemented in CosmoMC [43] and the nested sampling algorithm of MultiNest [44][45]. We use a modified version of the CAMB [46] code in order to compute the CMB anisotropies spectrum for different values of the parameters describing the step-like inflationary model, as in Eqs. (2) and (3). The Gelman and Rubin criteria [51][52] is used to evaluate the convergence of the Monte Carlo Markov chain (MCMC) analysis, demanding that \( R - 1 \leq 0.02 \). In our Bayesian analysis we use the most accurate Importance Nested Sampling (INS) [46][53] instead of the vanilla Nested Sampling (NS), requiring INS Global Log-Evidence error < 0.1.

A. Priors

As mentioned earlier, the class of models considered in this analysis is able to produce localised oscillations in the primordial power-law potential. The oscillation spot is set by the parameter \( \eta_f \) in the Eqs. (2) and (3) and depends on the scale where the wavelengths cross the horizon. We note that, for increasing values of the step-crossing time parameter, the oscillation is shifted to larger scales (lower multipoles), until it disappears completely from the TT spectrum for values of \( \ln(\eta_f/\text{Mpc}) \) > 10. At the same time, for values of \( \ln(\eta_f/\text{Mpc}) \) < 5 the oscillation is shifted to the high-\( ℓ \) part of the spectrum, i.e., to scales without interest to the present study.

Previous works [20][55] identified two different ranges of multipoles for these oscillations from the data: the first one lies in the interval 10 < \( ℓ \) < 60 (hereafter Oscill-1) whereas the second one lies in the interval 150 < \( ℓ \) < 300 (hereafter Oscill-2). In particular, these oscillation ranges are distinguished by different priors on the \( A_f \) parameter, i.e. \( A_f > 0.5 \) for Oscill-1 and \( A_f < 0.5 \) for the Oscill-2. We run our preliminary test following the guidelines of the previous studies, i.e., as-

| Parameter | Prior (Oscill-1) | Prior (Oscill-2) |
|-----------|-----------------|-----------------|
| \( A_f \) | Uniform(0.5, 1) | Uniform(0, 0.2) |
| \( \ln η_f \) | Uniform(7, 8) | Uniform(6.5, 8) |
| \( \ln x_d \) | Uniform(0, 1) | Uniform(3, 5) |

TABLE I: Priors on the features parameters of Eqs. (2) and (3).
assuming the same priors on the feature parameter $A_f$, and setting for the parameters $\ln(n_f/Mpc)$ and $\ln x_d$ the intervals $[5:10]$ and $[-1:5]$, respectively. Using the CosmoMC we found the value where the step parameter posterior probability drops to zero, and we set the priors for our analysis as shown in Table I. In order to estimate the impact of the prior choice on the $A_f$ parameter, we also repeat the analysis using the relaxed priors $A_f [0:1]$, $\ln(n_f/Mpc) [5.5:8.5]$, $\ln x_d [0:2]$.

It is important to highlight the central role played, in the CosmoMC analysis, by the prior ranges choice, since the introduced $n_f$ and $x_d$ step parameters show multimodal posterior probability distributions. The MCMC analysis, indeed, can fail to fully explore all peaks which contain significant probability, especially if the peaks are very narrow. The MULTINEST algorithm, on the contrary, do not shows the same analysis problem. For this reason we use the CosmoMC code only for the preliminary parameters estimation and we present in this work only the MULTINEST analysis results.

B. Data sets

We use the second release of Planck data \cite{2,34} (hereafter TT+lowP), namely the high-$\ell$ Planck temperature data (in the range of $30 < \ell < 2508$) from the 100-,143-, and 217-GHz half-mission TT cross-spectra and the low-P data by the joint TT, EE, BB and TE likelihood (in the range of $2 < \ell < 29$). More precisely, the latter data come from the best-fit temperature map obtained by the commander component separation algorithm applied to Planck 30-857 GHz data, jointly with the Wilkinson Microwave Anisotropy Probe (WMAP) 9-year observations between 23 and 94 GHz \cite{35} and the Haslam et al. 408 MHz survey \cite{36}; the E and B maps are obtained from the 70 GHz maps using their 30 and 353 GHz maps as foreground templates.

We combine the CMB data with the matter power spectrum measurements from the Baryon Oscillation Spectroscopic Survey (BOSS) CMASS Data Release-11 sample (covering the redshift range $z = 0.43 - 0.7$) of the Sloan Digital Sky Survey (DR11-SDSS) experiment (hereafter “CMB+SDSS”). We use the data sets of \cite{37} and publicly available in the SDSS Collaboration website (www.sdss3.org).

The best-fit CMB angular spectra and the matter power spectra for Oscill-1 and Oscill-2 models are shown in Fig. 1 and Fig. 2. In order to obtain the results we use binned high-$\ell$ TT data for the analysis of the Oscill-1 model since this type of oscillation is placed in the low-$\ell$ region of the spectrum, which means that details on the high-$\ell$ part do not add any information. On the other hand, for the Oscill-2 model the oscillations are located at small scales and show higher frequency (see Fig. 1 right panel). Therefore, for the analysis of this latter model we use the unbinned version of the high-$\ell$ TT data in order to maximise the sensitivity of sharp features that lie inside the single bin.

C. Statistical Model Selection

In order to probe possible features in the primordial inflationary power spectrum, we perform a Bayesian model comparison considering three models, namely, the standard $\Lambda$CDM scenario and the two models with localised oscillations in the power-law potential named earlier as Oscill-1 and Oscill-2. In this kind of analysis the “best” model is the one that achieves the best compromise between quality of fit and predictivity. Indeed, while a model with more free parameters will always fit the data better (or at least as good as) a model with less parameters, such added complexity ought to be avoided whenever a simpler model provides an adequate description of the observations. The Bayesian model comparison offers a formal way to evaluate whether the extra complexity of a model is required by the data, preferring the model that describes the data well over a large fraction of their prior volume. In this section, we briefly review the model comparison and introduce our notation (we refer the reader to \cite{45,50,51} for some recent applications of Bayesian model selection in cosmology).

Let us consider two competing models, $M_i$ and $M_j$, whose $n_j$ parameters (of model $M_j$) are common to $M_i$, which in turn has $n_i - n_j$ extra parameters. The posterior probability for the parameters vector $\theta$ (of length $n_i$) given the data $x$ under the model $M_i$ comes from Bayes’ theorem:

$$p(\theta|x, M_i) = \frac{p(x|\theta, M_i)\pi(\theta|M_i)}{p(x|M_i)} ,$$

and similarly for $M_j$. The term $p(x|\theta, M_i)$ is the likelihood, while the $\pi(\theta|M_i)$ the prior probability distribution function. The normalization constant in the denominator is called evidence $E$, and is the marginal likelihood for the model $M_j$:

$$\pi(\theta|M_i) \equiv E_{M_i} = \int d\theta p(x|\theta, M_i)\pi(\theta|M_i) .$$

The posterior probability of the $M_i$ model given the data is written as

$$p(M_i|x) \propto E_{M_i}\pi(M_i) .$$

Assuming no a priori preference about any model ($\pi(M_j) = \pi(M_i)$), the ratio of the posterior probabilities of the two models (the so-called Bayes Factor) is given by

$$B_{ij} = \frac{E_{M_i}}{E_{M_j}} .$$

The more complex model $M_i$ will (if $M_j$ is nested) inevitably lead to a higher (or at least equal) likelihood, but the evidence will favor the simplest model if the fit
TABLE II: 68% confidence limits for the cosmological and step parameters. The first columns-block refer to a minimal \( \Lambda \)CDM model with a featureless spectrum, using binned high-\( \ell \) TT data; the second columns-block show the constraint of the inflationary step-like model using Oscill-1 prior of Tab.1 and binned high-\( \ell \) TT data; the last columns-block show the constraint of the inflationary step-like model using Oscill-2 prior of Tab.1 and unbinned high-\( \ell \) TT data. The \( \Delta \chi^2_{best} \) and the \( \ln B_{M',i} \) of Oscill-2 analysis refers to the difference with respect to the \( \Lambda \)CDM analysis using the unbinned high-\( \ell \) TT data. These results refer to the Bayesian analysis results of the MultiNest code analysis.

| Parameter | \( \Lambda \)CDM model | Oscillation-1 | Oscillation-2 |
|-----------|-----------------------|--------------|--------------|
|           | TT+lowP               | CMB+SDSS     | TT+lowP       | CMB+SDSS     | TT+lowP       | CMB+SDSS     |
| 100\( \Omega \)kms\( ^2 \)Mpc\(^{-1} \) | 2.225 ± 0.02         | 2.232 ± 0.02 | 2.222 ± 0.021| 2.229 ± 0.02 | 2.225 ± 0.021| 2.233 ± 0.02 |
| \( \Omega \)\( \Lambda \) | 0.1194 ± 0.0019       | 0.1182 ± 0.0017| 0.1202 ± 0.0019| 0.1188 ± 0.0017| 0.1195 ± 0.0019| 0.1183 ± 0.0017|
| 100\( \theta \) | 1.04091 ± 0.00043 | 1.04104 ± 0.00042 | 1.04108 ± 0.00045 | 1.04097 ± 0.00042 | 1.04094 ± 0.00043 | 1.04111 ± 0.00040 |
| \( \tau \) | 0.083 ± 0.008 | 0.084 ± 0.008 | 0.083 ± 0.008 | 0.084 ± 0.008 | 0.083 ± 0.008 | 0.083 ± 0.008 |
| \( r_s \) | 0.9664 ± 0.0052 | 0.9685 ± 0.0050 | 0.9637 ± 0.0052 | 0.9660 ± 0.0050 | 0.9667 ± 0.0053 | 0.9685 ± 0.0049 |
| \( \ln 10^{10} A_s \) | 3.100 ± 0.0162 | 3.099 ± 0.0166 | 3.102 ± 0.0164 | 3.100 ± 0.0166 | 3.100 ± 0.0163 | 3.097 ± 0.0170 |
| \( \ln \eta_f /Mpc \) | – | – | – | – | 7.18 ± 0.10 | 7.18 ± 0.09 |
| \( \ln x_d \) | – | – | – | – | 0.49 ± 0.27 | 0.49 ± 0.27 |
| \( \sigma_8 \) | 0.835 ± 0.009 | 0.829 ± 0.009 | 0.836 ± 0.009 | 0.831 ± 0.009 | 0.834 ± 0.009 | 0.828 ± 0.009 |
| \( H_0 \) | 67.48 ± 0.83 | 68.00 ± 0.74 | 67.10 ± 0.85 | 67.72 ± 0.73 | 67.42 ± 0.85 | 67.98 ± 0.73 |

\( ^a \kappa_0 = 0.05\text{Mpc}^{-1} \)

\( ^b \)\( [\text{km} \text{s}^{-1} \text{Mpc}^{-1}] \)

\( ^c \)The associated error is calculated with the simple error propagation formula, assuming that the two measurements are uncorrelated: \( \sigma^2 (\ln B_{ij}) = \sigma^2 (\ln E_1) + \sigma^2 (\ln E_j) \)

\( \Delta \chi^2_{best} \) and \( \ln B_{M',i} \)

\[ \Delta \chi^2_{best} = \frac{1}{2} \sum_{ij} B_{ij} \left( \Delta \theta_i \cdot \Delta \theta_j - \Delta \theta_{ij} \right) \]

\[ B_{ij} = \frac{d\theta p(x|\theta, M_i) (\Delta \theta_i \cdot \Delta \theta_{ij})}{d\theta p(x|\theta, M_j) (\Delta \theta_j \cdot \Delta \theta_{ij})} \]  

\text{FIG. 1: Temperature power spectrum for the two inflationary step-like models best-fit values in comparison with the \( \Lambda \)CDM model best-fit (black line). Left: best-fit values for the Oscill-1 model using TT+lowP data. In the small box a zoom at 5 < \( \ell \) < 80. Right: best-fit values for the Oscill-2 model using CMB+SDSS data. In the small box a zoom at 150 < \( \ell \) < 280 is nearly as good, through the smaller prior volume. We assume uniform (and hence separable) priors in each parameter, such that \( \pi(\theta|M_i) = (\Delta \theta_i \cdot \Delta \theta_{ij})^{-1} \) and \( B_{ij} = \frac{d\theta p(x|\theta, M_i)}{d\theta p(x|\theta, M_j)} \).

\text{In order to rank the models of interest, we adopted the following scale to interpret the values of } \ln B_{ij} \text{ in terms of the strength of the evidence of a chosen reference model } (M_j): \ln B_{ij} = 0 - 1, \ln B_{ij} = 1 - 2.5, \ln B_{ij} = 2.5 - 5, \text{ and } \ln B_{ij} > 5 \text{ indicate, respectively, an inconclusive, weak, moderate and strong preference of the model } M_i \text{ with respect to the model } M_j. \text{ Note that negative values of } \ln B_{ij} \text{ means support in favour of the model } M_j. \text{ We refer to } [19] \text{ for a more complete discussion about this scale, that is a revised and more conservative version of the so-called Jeffreys’ scale} [57].
FIG. 2: Matter power spectrum for the best-fit ΛCDM model (dotted black line) and for the two step-like models: Oscill-1 best-fit values using TT+lowP data (red line), Oscill-2 best-fit values using CMB+SDSS data (blue line).

III. RESULTS

The main quantitative results of our analysis are shown in Tab. 1. Firstly, we assume the minimal ΛCDM model and use the CMB and CMB+SDSS data sets discussed earlier. From Fig. 1 one can see that the addition of the galaxy data (solid line) shows a preference for lower values of Ω_c h^2 and for higher values of n_s, with respect to the analysis using only CMB (TT+lowP) data. From Tab. 1 we also note a slightly improvement of the constraints on the Hubble parameter H_0 and a preference for higher value of σ_8. A good concordance between the results using binned and unbinned high-ℓ TT data is verified. For brevity, however, we report in Tab. 1 only the ΛCDM analysis using the binned high-ℓ TT data. It is also worth mentioning that the values of Δχ^2_{best} and ln B_{ij} for the Oscill-2 model are obtained with respect to the ΛCDM analysis using the unbinned high-ℓ TT data.

For the central mean values given in Tab. 1, we show in Fig. 3 the primordial scalar power spectrum for the two step models. We note that the two primordial oscillations start around the same scale (approximately the same value of η_f) but show very different amplitude and damping properties, which in turn produce very different features. The primordial power spectrum of the Oscill-1 model (red line) has a high amplitude and damping parameters values, and produces features at the scale interval 10^{-3} \lesssim k \lesssim 10^{-2}. In principle, this makes possible to detect them using only CMB data, since the current LSS data cover scales of k \gtrsim 10^{-2}. [37].

In Fig. 4 we show the posterior probability distribution for the step parameters values. As expected, the addition of the large-scale structure data improves the constraints on these parameters, mainly those related to the Oscill-2 model whose features extends to lower scales (see Fig. 1). In comparison with the TT+lowP constraints (blue dashed line), we note the higher constraints on the frequency parameter ln η_f/Mpc, which for the CMB+SDSS data shows a bimodal distribution.

Table 1 also shows that the constraints on the usual cosmological model parameters are not significantly affected by the presence of primordial features. Moreover, in the case of the Oscill-1 model, our bounds on the step parameters are consistent with previous analysis [20] using the first Planck release (2013) with the only exception of the amplitude parameter A_f, which now prefers lower values. On the other hand, our results for the TT+lowP data also shows a moderate evidence (ln B_{ij} = 3.91±0.03) in favor of the Oscill-1 model with respect to the ΛCDM scenario. If we relax the step-parameter priors to

A_f [0 : 1], \ ln(η_f/Mpc) [5.5 : 8.5], \ ln x_d [0 : 2], (9)

the moderate evidence of Oscill-1 becomes weak, and the distribution of probability shows a weak secondary peak, which extends the explored parameter space. Our results seem to be in disagreement with those of the Planck Collaboration [2], where a featureless primordial potential is preferable over the step-like models. Actually, the constrained oscillation in Ref. [2] refers to a different parametrization model [21] and shows different step parameters values, whose overall effect is to produce smaller and deeper oscillations with respect to the results of this work, as shown in Fig. 4. At the same time, the parametrization used in this work produce a different behaviour in the scales immediately before where the oscillation occurs.

For the Oscill-2 model, our analysis shows an a inconclusive evidence (ln B_{ij} = 0.58 ± 0.04) relative to the ΛCDM scenario using the TT+lowP dataset. We also observe that the addition of P(k) data weakens the evidence of Oscill-1 model relative to the standard cosmology and strengthens the evidence of this latter with respect to the Oscill-2 model.

IV. CONCLUSION

We have performed a Bayesian model selection statistics to compare the observational viability of a class of inflationary models with step-like features in the inflaton potential and the ΛCDM cosmology using the most up-to-date CMB and LSS datasets. The step-like inflationary potentials studied are able to produce features in the primordial scalar power spectra, inducing an oscillation in the anisotropy power spectrum with magnitude, extent and position which depend on three step parameters. We have considered two types of models beyond the minimal ΛCDM model: Oscill-1 model, whose features lie in the multipole range 10 < ℓ < 60, and the Oscill-2 model, which produces features at multipole 150 < ℓ < 300 (see Fig. 1).

In order to perform our analysis, we have used an approximate form of the power spectrum, as given in Eqs.
FIG. 3: 68% and 95% confidence regions on ΛCDM model for Planck TT+lowP data (black dotted line) and CMB+SDSS data (black solid line). The numerical results of these analyses are reported in the first columns-block of Tab. II.

FIG. 4: Primordial power spectrum for the two step models: Oscill-1 (red line) with $A_f = 0.75$, $\ln(\eta/\text{Mpc}) = 7.18$ and $\ln x_d = 0.49$ and Oscill-2 (blue line) with $A_f = 0.04$, $\ln(\eta/\text{Mpc}) = 7.21$, and $\ln x_d = 4$. The $P(k)$ measurements from the DR11 of the SDSS Collaboration. For the ΛCDM model, our analysis shows a good concordance between the results using the CMB data only and the extended dataset (CMB+SDSS). As the main result of this analysis, we have shown that the Oscill-1 model provides a better fit to current CMB data than does the standard ΛCDM model, with a moderate Bayesian evidence in favor of the step-like potential (Oscill-1). Such result, however, seems to be in disagreement with the one reported by the Planck Collaboration, in which a featureless primordial potential is preferable over the step-like models. When the extended data set (CMB+SDSS) is considered, the evidence of the Oscill-1 model relative to the standard cosmology becomes weak, just as the Bayesian evidence of the ΛCDM cosmology with respect to the Oscill-2 model.

Finally, as shown in Fig. 4, our best-fit scenario (Oscill-1 model) deviates significantly from the standard model only at very low values of $k$. We, therefore, expect to
FIG. 5: One-dimensional posterior probability densities. Top: Oscill-1 step parameters analysis reported in Tab. II central columns-block, using TT+lowP data (red dashed line) and CMB+SDSS data (red solid line). Bottom: Oscill-2 step parameters analysis reported in the last columns-block of Tab. II using TT+lowP data (blue dashed line) and CMB+SDSS data (blue line).

FIG. 6: Primordial power spectrum for the Oscill-1 best-fit model (red line) with $A_f = 0.728$, $\ln(\eta/Mpc) = 7.20$ and $\ln x_d = 0.75$, and for the parameterisation discussed in the Planck-2015 analysis [2] (black line) with the best-fit values $A_f = 0.347$, $\ln(\eta/Mpc) = -3.10$, and $\ln x_d = 0.342$.

verify the reality of these features in the CMB and matter power spectra with data from the future release of the Planck collaboration, as well as from the next generation of very deep galaxy surveys like, for example, the Dark Energy Spectroscopic Instrument (DESI) [59] and the Javalambre Physics of the Accelerating Universe Astrophysical Survey (J-PAS) [60].

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