Four winters of photometry with ASTEP South at Dome C, Antarctica

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ABSTRACT

Context. Dome C in Antarctica is a promising site for photometric observations thanks to the continuous night during the Antarctic winter and favorable weather conditions.

Aims. We developed instruments to assess the quality of this site for photometry in the visible and to detect and characterize variable objects through the Antarctic Search for Transiting ExoPlanets (ASTEP) project.

Methods. Here, we present the full analysis of four winters of data collected with ASTEP South, a 10 cm refractor pointing continuously toward the celestial south pole.

Results. We achieved nearly continuous observations over the winters. We improved the instrument over the years and developed specific data reduction methods. We measure an average sky background of 20 mag arcsec$^{-2}$ in the 579–642 nm bandpass. We built the lightcurves of 6000 stars and developed a model to infer the photometric quality of Dome C from the lightcurves themselves. The weather is photometric 67.1 ± 4.2 % of the time and veiled 21.8 ± 2.0 % of the time. The remaining time corresponds to poor quality data or winter storms. We analyzed the lightcurves of ζ Oct and HD 184465 and find that the amplitude of their main frequency varies by a factor of 3.5 and 6.7 over the four years, respectively. We also identify 34 new variable stars and eight new eclipsing binaries with periods ranging from 0.17 to 81 days.

Conclusions. The phase coverage that we achieved with ASTEP South is exceptional for a ground-based instrument and the data quality enables the detection and study of variable objects. These results demonstrate the high quality of Dome C for photometry in the visible and for time series observations in general.

Key words. Methods: observational – Methods: data analysis – Techniques: photometric – Site testing – Stars: variables: delta Scuti – Binaries: eclipsing

1. Introduction

Dome C in Antarctica offers exceptional conditions for time-series observations thanks to the continuous night during the Antarctic winter and very favorable atmospheric conditions. Dome C is located on a summit of the high Antarctic plateau at an altitude of 3233 meters, 1100 km away from the coast, at coordinates 75°06′S – 123°21′E. The site is equipped with the French-Italian Concordia station, which was built between 1999 and 2005 and runs all year long since 2005. The station is designed to host various scientific experiments, and can host about 80 people in summer and 15 people for winter-overs. Site testing for astronomy began in the early 2000s and revealed exceptional conditions: a clear sky most of the time including during the winter, low wind speeds, a dry atmosphere, and low scintillation (Aristidi et al. 2003, 2005; Lawrence et al. 2004; Ashley et al. 2005; Geissler & Masiadi 2006; Kenyon et al. 2006). Experiments also found a free-atmosphere seeing above a boundary layer of about 30 meters (Agabi et al. 2006; Trinquet et al. 2008; Aristidi et al. 2009). Night-time astronomical observations started in 2006. They revealed a high duty cycle for spectroscopy and showed that high quality time-series photometry was achievable from this site (Mosser & Aristidi 2007; Strassmeier et al. 2008; Briguglio et al. 2009). The continuous coverage during the winter is beneficial to the search for and study of variable stars and transiting exoplanets.

We developed the Antarctic Search for Transiting Exoplanets (ASTEP) project to search for and characterize transiting exoplanets and to determine the quality of Dome C for photometry in the visible (Fressin et al. 2005). The project started in 2005 and two instruments were installed at Dome C. The main instrument, ASTEP 400, is a 40 cm Newton telescope designed and built by our team to perform high precision photometry under the extreme conditions of the Antarctic winter (Daban et al. 2010). ASTEP 400 was installed in 2010 and observed nomi-
nally over four winters. The results are presented in Abe et al. (2013); Mékarnia et al. (2016); Crouzet et al. (2016); Chapellier et al. (2016) and the lightcurves of more than 300,000 stars are publicly available. The 2017 and 2018 winter campaigns were dedicated to the photometric monitoring of β Pic during the transit of the Hill sphere of its planet β Pic b (Mékarnia et al. 2017).

A first instrument, ASTEP South, was installed in 2008 to acquire preliminary data for the project, to assess its feasibility, and to identify the challenges of photometric observations from Dome C. ASTEP South observed over four winters from 2008 to 2011. The results from a preliminary analysis of the first winter are reported in Crouzet et al. (2010). In this paper, we present the full analysis of the ASTEP South data.

2. Observations

ASTEP South consists of a 10 cm refractor, a front-illuminated 4096 × 4096 pixel CCD camera, and a simple mount in a thermally isolated enclosure. The instrument is completely fixed and points continuously toward the celestial south pole during the Antarctic winter, which minimizes jitter noise and airmass variations during the observations. This setup leads to stars moving circularly on the CCD with a one-sidereal day period and to elongated point spread functions (PSF). The field of view is 3.88 × 3.88 arcmin², corresponding to a pixel size of 3.41 arcsec on the sky. The exposure time is 29 s with 10 s overheads between each exposure. The PSF full width half maximum (FWHM) is nominally 2 pixels, not taking into account the elongation due to rotation. At the border of the field, the rotation rate of the stars on the CCD is 0.15 pixel s⁻¹, corresponding to a PSF elongation of 4.33 pixel during one exposure. The observing bandpass is equivalent to a large R band (600–900 nm). The enclosure is thermalized to −20°C and fans are mounted inside to homogenize the temperature. The enclosure is closed by a double glass window to minimize temperature fluctuations (the outside temperature varies between −50 and −80°C during the winter at Dome C). Software programs were developed by our team to control the camera, run the acquisitions, transfer and save the data. This instrumental setup is presented in more detail in Crouzet et al. (2010). ASTEP South was installed at the Concordia station in January 2008 and observed during the Antarctic winters 2008, 2009, 2010, and 2011. Spurious reflections presumably due to moonlight were identified during the Antarctic winters 2008, 2009, 2010, and 2011. The 2008 and 2010 observations were analyzed in Crouzet et al. (2010). The results from a preliminary analysis of the first winter are presented in more detail in Crouzet et al. (2010). In this paper, we present the full analysis of the ASTEP South data.

Table 1: Number of science images and amount of data acquired by ASTEP South. Calibration frames are not included.

| Observing season | Number of science frames | Total size [Terabytes] |
|------------------|--------------------------|------------------------|
| 2008             | 125,003                  | 4.0                    |
| 2009             | 223,756                  | 7.1                    |
| 2010             | 241,026                  | 7.7                    |
| 2011             | 279,295                  | 8.9                    |
| Total            | 869,080                  | 27.8                   |

3. Data reduction

3.1. Star catalog

The ASTEP South target star catalog is a combination of UCAC3, UCAC2, TYCHO-2, and APASS. These catalogs were obtained from VizieR at the Centre de Données Astronomiques de Strasbourg² or from the American Association of Variable Star Observers Photometric All-Sky Survey website³, via requests centered on the celestial south pole (RA = 0°, Dec = −90°) with a search radius of 2°. Our main catalog is UCAC3. Most of the bright stars are flagged as unreliable data; we use the TYCHO-2 entries for these stars. TYCHO-2 does not add any stars. We also noticed that a small fraction of stars in our images were missing in UCAC3 but were present in UCAC2; we use the UCAC2 entries for these stars. Finally, APASS is used to gather complementary information on the magnitudes of the targets. We keep all the magnitude and flag information available in these catalogs. For multiple entries, defined as having a separation of less than 7 arcsec (2 pixels in the ASTEP South images), only the brightest star is kept; the other ones may be catalog artifacts or stars that would not be resolved by ASTEP South. We use the UCAC3 fmag magnitude defined in the bandpass 579–642 nm as our reference magnitude, the UCmag defined in the same bandpass for stars present only in UCAC2, or the Sloan i' magnitude for stars present only in APASS.

Due to strong image vignetting (~25% at the edges) and to PSF elongation, the limit magnitude of stars that are visible in the ASTEP South images is lower at the border than at the center of the field. The limit flux $F_{\text{lim}}$ is well represented by a linear function after a certain distance from the center:

$$F_{\text{lim}} = \begin{cases} F_{\text{lim,c}} & \text{if } r \leq r_0 \\ F_{\text{lim,c}} (1 - a (r - r_0)) & \text{if } r > r_0 \end{cases}$$ (1)

where $F_{\text{lim,c}}$ is the limit flux at the center of the image, $r$ is the radial distance from the center in pixels, $r_0 = 300$ pixels, and $a = 4.284 \times 10^{-4}$ pixel⁻¹. This corresponds to a 1.5 magnitude difference between the center and the edges, which have a limit magnitude $f_{\text{mg}}^{\text{lim}}$ of 16 and 14.5 respectively. We used this function to select only stars below the limit magnitude after calculating their nominal radial position in the images.

In total, the ASTEP South catalog comprises 5954 target stars including 5830 from UCAC3, 91 from TYCHO-2 (almost all the stars with fmag < 9.8), 29 from UCAC2, and four from APASS. A visual inspection of these targets superimposed on a high quality ASTEP South image shows that this catalog is nearly complete for our purpose.

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¹ https://astep.vo.oca.eu/
² http://vizier.u-strasbg.fr/
³ http://www.aavso.org/apass
3.2. Image calibration

Calibration images were taken during the observations, with a sequence of one bias between every science image, and one dark every ten science images. Although very consuming in terms of disk space, the large number of bias was useful for a sanity check, for example to measure the bias level as a function of the camera's electronics temperature. We removed most of the oscillations shown in Fig. 10 of Crouzet et al. (2010) by improving the thermalization after the first campaign, but non-periodic temperature fluctuations remain. To correct for this, we computed the median value of each bias and interpolated them linearly to obtain the bias level for each science image. A “masterbias” and a “masterdark” were created for each day by stacking images together with successive median filters. The dark current itself is negligible and was not detected. The median level of the darks is lower than that of the bias by about 50 analog to digital units (ADU) for an unexplained reason, and the median level of the science images matches that of the darks.

Around noon local time, the Sun is high enough (below the horizon) to slightly increase the sky background (Crouzet et al. 2010). Although not detectable by naked eye, this illuminates enough the CCD to create “twilight” images, which we used to build flat-fields. We used images with median values between 8,000 and 30,000 ADU (the pixel dynamic range is 65,535 ADU). Between a few tens and two hundred such images were recorded daily except in June where the Sun remains too low. Thanks to their motion around the CCD, stars could be removed by stacking images together and applying successive medians. We built one flat-field for each day with the following procedure: bright images were calibrated from bias and dark, corrected from hot pixels, normalized by their median in the 512×512 pixel central region, and divided into sets of 15 images. We interleaved the different sets instead of grouping consecutive images, so that images in each set were stacked together with a median filter. We obtained up to seven intermediate flat fields. We stacked them together with an outlier resistant mean to create one flat-field. We found that this method was significantly more efficient at removing stars than applying a unique median filter to all the images together. Finally, these daily flat-fields were stacked together with a median to create “masterflats” for each winter. This method yielded high quality masterflats, as shown in Fig. 1.

![Fig. 1: Masterflat (left) and same image after subtracting the low frequency component (right) for 2009. The vignetting is ~25%. The color scale is arbitrary.](image)

Random failures of the camera shutter began after two winter operations of temperature at +5°C, corresponding to ~393,000 open-close cycles. This issue became more severe in 2010 and 2011. We increased the temperature to +30°C around the shutter for the 2011 campaign, which improved its reliability but did not solve the issue completely. Science images, darks, and bias were affected (Fig. 2). We developed a calibration method to correct for these effects. The masterbias and masterdark did not provide a satisfactory calibration, so we built a third type of calibration image. First, we cleaned the masterbias from residual light: we removed a low frequency component from each row and replaced it by a linear function between both sides (this linear trend also appeared on bias unaffected by shutter issues). Then, we adjusted the median value of the masterbias and of the corrected masterbias to match that of the masterdark; the medians were calculated in 200 pixel wide boxes located at the corners. Finally, we stacked these three images together with a median to create a “mastermix” (Fig. 2). This median can be seen as a conditional statement: each part of the mastermix corresponds to one of these images. We found that it was a good representation of the shutter effects seen on the science images, and provided a satisfactory calibration.

Our science images were calibrated as:

\[
C = \frac{(I - M - b) \times g}{F}
\]

where \(C\) is the calibrated image, \(I\) is the raw image, \(M\) is the masterdark or the mastermix in case of shutter issues, \(b\) is the bias variations due to temperature fluctuations (taken around its median value for each day), \(g\) is the gain of the CCD, and \(F\) is the masterflat. We measured the gain during the characterization of the ASTEP cameras in the lab using the approach of Leach et al. (1980); Leach (1987). We took pairs of images at different illumination levels spanning the full dynamical range of the CCD. Then, the gain is given by \(g = 2M/V\), where \(V\) is the variance of the differential image and \(M\) is the mean of the image (we neglected the read-out noise in this calculation). The \(V - M\) relation was well represented by a linear function below 31,360 ADU and by a third order polynomial above. This yielded a constant gain of 2 e^-/ADU below 31,360 ADU and an increasing gain from 2 to 2.2 e^-/ADU between 31,360 and 60,000 ADU. Finally, we corrected for hot pixels. They were present in the masterdarks but not in the masterbias. Pixels in the masterdarks that exceeded their value in the masterbias by more than 150 ADU were flagged for all the images of that day and were replaced by the average of the four adjacent pixels. An example of calibrated image is shown in Fig. 3.

3.3. Field matching algorithm

In each science image, we identified point sources using the FIND procedure in IDL, with a threshold intensity of five times the sky standard deviation. These point sources were compared to the ASTEP South catalog using a home-made field matching algorithm. The star coordinates in the catalog (RA, Dec) are converted into image coordinates \((x, y)\). Then, we perform iterations.
on pairs of stars using the 20 brightest non saturated stars. For each pair in the ASTEP South image and each pair in the catalog, we calculate the geometrical transformation that matches one pair to the other, skipping iterations where both stars do not have approximately the same separation. We apply this transformation to the 400 brightest stars and evaluate the matching by a $\chi^2$, where $\chi$ is the median distance between a star in the transformed image and the closest star in the catalog. The best matching is found very quickly and yields a typical precision of 0.4 pixel. Then, $\chi^2$ is minimized using the simplex method in order to better estimate the transformation parameters. This yields a precision of 0.2 pixel. Finally, we calculate and correct for distortions using a regression in $y$ yields a precision of 0.2 pixel. Finally, we calculate and correct for distortions using a regression in $y$, $x$ of 0.4 pixel. Then, applying this transformation to the 400 brightest stars using an aperture radius of 5 pixels. For each star, we kept the aperture radius that yielded the smallest root mean square (RMS). In the case of neighboring stars, the resulting aperture was generally too large because it contained several stars, and the optimal aperture appeared as a local RMS minimum rather than as a global minimum. This was corrected by forcing the aperture radius to remain within ±2 pixels of the typical aperture for this star’s magnitude. In the end, each star had its own optimized aperture size, which remained the same for all the photometric reduction.

After reducing all the days of a winter, we created a single lightcurve file for each star with the flux measurements and information on the aperture position, centroid position, and local sky background. We also built a reference lightcurve using 60 bright stars with low RMS that we identified from one typical day. We normalized these 60 lightcurves by their respective medians, and averaged them using an outlier resistant mean with a 3$\sigma$ cut-off. The resulting reference lightcurve was also normalized by its median.

3.4. Photometric algorithm

The ASTEP South PSFs are particularly challenging for high-precision photometry. We developed our own photometric algorithm to take into account the elongated PSFs, and their elongation directions and positions that vary in time as the stars rotate on the CCD. The aim was to maximize the measured flux while reducing the sky background, readout noise, hot pixels, cosmic rays, and contamination by surrounding stars compared to circular apertures.

We computed the apertures for each star in each image. The apertures consist of two half-circles joined by a rectangle which length and orientation depend on the position of the star relative to the celestial south pole (Fig. 4). The aperture is quasi-circular for stars close to the south pole and is the most elongated at the edges of the field. To measure the stellar flux, we subdivide each pixel within the aperture into a grid of 11 × 11 subpixels, project the aperture onto this grid, and sum the flux at the subpixel level; this is equivalent to weighting each pixel by the fraction of its area within the aperture. The sky background is calculated in an annulus of 11 pixels starting 5 pixels outside the star aperture, multiplied by the aperture area, and subtracted to the measured flux.

To calculate the optimal aperture size for each star, defined by the radius of the half-circles, we ran the photometric algorithm on a typical day with radii ranging from 0.5 to 10 pixels by steps of 0.05. The lightcurves were calibrated by a reference lightcurve, calculated from the median of the 30 non saturated brightest stars using an aperture radius of 5 pixels. For each star, we kept the aperture radius that yielded the smallest root mean square (RMS). In the case of neighboring stars, the resulting aperture was generally too large because it contained several stars, and the optimal aperture appeared as a local RMS minimum rather than as a global minimum. This was corrected by forcing the aperture radius to remain within ±2 pixels of the typical aperture for this star’s magnitude. In the end, each star had its own optimized aperture size, which remained the same for all the photometric reduction.

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3.5. Correction of systematic effects

We divided the target stars’ lightcurves by the reference lightcurve and binned them per 400 second intervals (6.7 minutes, approximately ten points). Then, we identified systematic variations and applied several corrections. The largest variations were due to the rotation of stars on the CCD and were on the order of a few percent, sometimes up to 10%. They formed a pattern with a one-sidereal day period, which was different for each star and was usually repeatable. However, several complications arose. First, we sometimes modified the pointing during the winter, for example after improving the thermal control or after other interventions on the instrument; this changed the one-day patterns. Second, each pattern was split into two or more patterns that are shifted in flux, corresponding to different states of the FWHM (see Sect. 4.2). Third, we noticed flat-field variations over the winter possibly due to mechanical deformations and temperature variations inside the box, and we built a flat-field for each image calibration method (with or without shutter correction). In the end, we used between one and four flat-fields to calculate the optimal aperture size for each star, defined by the radius of the half-circles, we ran the photometric algorithm on a typical day with radii ranging from 0.5 to 10 pixels by steps of 0.05. The lightcurves were calibrated by a reference lightcurve, calculated from the median of the 30 non saturated brightest stars using an aperture radius of 5 pixels. For each star, we kept the aperture radius that yielded the smallest root mean square (RMS). In the case of neighboring stars, the resulting aperture was generally too large because it contained several stars, and the optimal aperture appeared as a local RMS minimum rather than as a global minimum. This was corrected by forcing the aperture radius to remain within ±2 pixels of the typical aperture for this star’s magnitude. In the end, each star had its own optimized aperture size, which remained the same for all the photometric reduction.

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over the winter. These pointing, FWHM, and flat-field effects were mixed together and yielded variable one-sidereal day patterns. To recover them, we split the data into sets that were as consistent as possible according to these parameters; in practice, each set corresponds to a specific time interval during the winter. Then, we recovered the patterns for each star and each set individually: we fold the star lightcurve over a given time interval at a one-sidereal day period, split it into 100 bins, calculate the median flux for each bin, interpolate them linearly, and divide the lightcurve by this function (Fig. 5). We applied this procedure twice. In some cases, because of poor weather, large FWHM variations, or small data sets, the patterns could not be recovered. The 2008 and 2011 data were divided into one and four sets respectively and were well corrected. The 2009 and 2010 data were much less stable and were divided into 12 and 16 sets respectively; some of them were completely discarded, whereas several large sets comprised most of the good data and were well corrected.

We converted fluxes into magnitudes and normalized them by their mean magnitude. From this point, we worked with the magnitude residuals. Then, we calibrated each star by a set of reference stars chosen as follows. We divided each lightcurve by all others and sorted the resulting lightcurves by their mean absolute deviation (MAD). The stars yielding the lowest MADs were used as reference stars and their lightcurves were averaged to build a reference lightcurve. We repeated this procedure for each star (so in the end each star had its own set of reference stars). We used three reference stars for the 600 brightest stars (fmag < 11.9) and ten reference stars for fainter stars.

We corrected for remaining systematic variations. We gathered the lightcurves in a star, epoch array with stars sorted by ascending magnitudes. Systematic trends appeared along the epochs and were common to stars of similar magnitudes. We corrected for them on a row by row (epoch by epoch) basis by computing a median smoothing over 200 stars and subtracting it to the initial row. This method yielded a better correction than Systematic Removal (SYSREM, Tamuz et al. 2005) for these data at this stage of the pipeline, because it is insensitive to variable stars, which are often interpreted as systematic effects by SYSREM. Our method provided more uniform lightcurves overall. After these corrections, we computed the average standard deviation for each epoch \( \sigma_e \) using stars of magnitudes between 12 and 13.5.

3.6. Data selection and final lightcurves

In the process, we flagged outliers departing by more than 4\( \sigma_e \) from each lightcurve, where \( \sigma_e \) is the standard deviation of the lightcurve. We flagged entirely the epochs with more than 50% of outliers, as well as the epochs with less than 500 detected stars. Considering individual stars, we flagged entirely the days with more than 50% of outliers, because the remaining points were generally off the regular one-sidereal day pattern.

We flagged the epochs with a median sky background greater than 500 ADU. We flagged the epochs with \( \sigma_p > 2\% \) because they correspond to the tail of the \( \sigma_e \) distribution (see Sect. 5). For 2009, a limit at 3\% would be more appropriate considering all the data; however the distribution is bi-modal, one mode corresponding to a nominal FHWM (~2 pixels) and the other to a larger FHWM (~3 pixels). To keep only nominal data and for consistency with the other years, we kept the limit \( \sigma_p < 2\% \). We also flagged the epochs at which more than 30% of the 1000 brightest stars had bad data (these epochs were often of poor quality for all the stars), and those at which more than 10% of

| \( R \) mag | 10 | 12 | 14 | 15.5 |
|----------|----|----|----|------|
| \( \sigma_e \) | 0.16 | 0.45 | 1.54 | 4.08 |
| \( \sigma_p \) | 0.23 | 0.55 | 1.79 | 4.14 |
| \( \sigma \) | 0.41 | 0.81 | 2.41 | 5.47 |

Table 2: RMS of the four-winter lightcurves over 6.7 minute bins for different magnitudes, in percent. \( \sigma_e \) is the theoretical RMS, \( \sigma_p \) is the point-to-point RMS, and \( \sigma \) is the true RMS. The 100 brightest stars had bad data (we found that this was a good proxy for data taken around full Moon). Finally, we flagged lonely epochs (epochs with less than 20% of good epochs within their 40 closest neighbours), which were often missed by other flagging constraints. The flagging information was used in the analysis of the Dome C photometric quality (Sect. 5) but the flagged epochs were discarded in the final lightcurves and in the search for variable objects (Sect. 7). The final lightcurves of the 5954 target stars are publicly available on the ASTEP lightcurve database website\(^4\). Examples of lightcurves of variable objects are shown in Fig. A.1.

3.7. Lightcurve precision

We calculated the precision of the four-winter lightcurves over 6.7 minute bins. We computed the true lightcurve RMS \( (\sigma) \) and the point-to-point RMS \( (\sigma_p) \) which does not take into account correlated noise. We also computed theoretical estimates of the stellar photon noise, readout noise, sky background noise, scintillation noise, and their quadratic sum, first considering circular PSFs and circular apertures, then taking into account the PSF elongation and specific aperture of each star \( (\sigma_t) \). The scintillation noise is obtained from Eq. 1 of Young (1967) (see also Dravins et al. 1998; Ryan & Sandler 1998). By analyzing turbulence profiles, Kenyon et al. (2006) suggest a potential gain of 3.6 at Dome C compared to Chilean sites. We did not apply this factor here: the scintillation noise remains well below other noise sources for ASTEP South, and using larger telescopes or studying a few bright stars would be necessary to confirm this gain. The lightcurve RMS as a function of stellar magnitude is shown in Fig. 6 and the average values for various magnitudes are reported in Table 2.

At the faint end, the point-to-point RMS is well explained by the theoretical RMS for elongated PSFs. At magnitudes 10 and 12, the point-to-point RMS is larger than the theoretical RMS by a factor 2.6, 1.8, 1.6, and 1.3 at magnitudes 10, 12, 14, and 15.5 respectively, indicating that additional sources of uncorrelated noise is present and affect mostly bright stars. The true RMS is larger than the theoretical RMS by a factor 2.6, 1.8, 1.6, and 1.3 at magnitudes 10, 12, 14, and 15.5 respectively, indicating that correlated noise also remains in the lightcurves. For most stars, the true RMS of the four-winter lightcurves is within a factor of two of the theoretical estimates. This could be explained by the particular issues related to the challenging observing mode of ASTEP South. Overall, this shows that good photometry over very long periods is achievable at Dome C.

4. Instrumental parameters and sky background

Besides stellar fluxes, we extracted several parameters in order to evaluate the functioning of the instrument and the quality of

\(^4\) https://astep-vo.oca.eu/
Dome C for photometry. We present the results in this section, which complements the analysis of Crouzet et al. (2010) for the first winter.

| Winter 2008  | Winter 2009  | Winter 2010  | Winter 2011  |
|-------------|-------------|-------------|-------------|
| TYC 9504-139-1 | TYC 9521-270-1 | TYC 9519-180-1 | TYC 9519-22-1 |

**Fig. 5:** Correction of the one-sidereal day variations. From left to right: winters 2008 to 2011. One star is given as an example for each winter; the stars’ IDs are indicated. The lightcurves are phase-folded at a one sidereal day period. First row: initial lightcurve. Second row: split into different data sets. Each set is represented by a color and the recovered pattern is shown as a line of the same color. Third row: lightcurve after division by the patterns. Fourth row: identification of outliers, indicated in red. Fifth row: lightcurve with outliers removed. Sixth row: final lightcurve after applying all the photometric correction and data selection steps.

4.1. **Pointing variations**

We investigated the mechanical stability of ASTEP South by measuring the motion of the celestial south pole on the CCD (Fig. 7). Within one day, the south pole is stable within one pixel: the standard deviation from its average position is typically 0.6
and 0.5 pixel in x and y, respectively. The long term motion over the winters is given in Table 3 and averages to 0.18 and 0.09 pixel per day in x and y, respectively. This long term motion is oriented close to the north-south axis which is the vertical axis of the instrument; this is evident in 2008 and 2010 and less evident in 2009 and 2011. This motion may correspond to a bending of the instrument under its own weight over time, modulated by thermal fluctuations inside the box affecting its deformation coefficients. A motion of the ice below the instrument is also possible. Overall, this motion is very small and would be completely negligible for any instrument equipped with a guiding system operating at Dome C.

| Year | $\Delta x$ [pixel unit] | $\Delta y$ [pixel unit] | $\Delta t$ [day] |
|------|-------------------------|-------------------------|------------------|
| 2008 | 25                      | 10                      | 128              |
| 2009 | 13                      | 8                       | 86               |
| 2010 | 25                      | 11                      | 115              |
| 2011 | 33                      | 16                      | 211              |

### 4.2. Point spread function variations

We measured the PSF variations using the 20 brightest stars in the 1000×1000 pixel central region of the image, where the elongation is negligible. We fitted the PSFs with a Gaussian function to derive the FWHM. This calculation was done by the day-to-day analysis software program and was used as an instrument control parameter. The nominal FWHM for ASTEP South is 2 pixels, and its variations over the four winters are shown in Fig. 8. Thermal stability is critical to a stable FWHM. For example, during the first part of the 2009 winter, the FWHM suffered extreme variations because of a fan failure. The recorded temperature inside the enclosure (around the probe) varied between −22 and −18°C but the actual variations and gradients were probably much stronger, inducing thermal deformations of the instrument. After restarting the fans, the recorded temperature remained stable within 0.2°C of the setting point and the FWHM remained stable between 2 and 2.5 pixels. Long-term variations (timescales longer than one day) in 2008 and 2010 are also correlated with thermal variations. Because any intervention on the instrument during the Antarctic winter is extremely delicate, the thermal control was not modified during these winters. Under stable thermal conditions, the FWHM recorded by ASTEP South is directly correlated with the atmospheric seeing, which yields short-term variations (timescales shorter than one day). This has been shown for the 2008 winter by Crouzet et al. (2010) and is also evident during the second half of the 2009 winter and during the 2011 winter, where increases of the FWHM correspond to poor seeing conditions.

As for any photometric instrument, periods affected by large FWHM variations result in poor photometry. In particular, FWHM variations are correlated with the one-sidereal day patterns, with different patterns for FWHM below and above a limit of either 2.5 or 3 pixels. Above this value, the pattern can be highly variable from day to day and was not always recovered, in particular in 2009 and 2010. These data were generally eliminated from the lightcurves.

### 4.3. Sky background

The sky background was computed in each image using the IDL SKY procedure, corrected for the zenith angle, converted from photo-electrons to instrumental magnitude, and calibrated to a $f_{\text{mag}}$ magnitude (579–642 nm bandpass). We performed this latter calibration by fitting a linear function between the stars’ instrumental magnitudes and their UCAC3 $f_{\text{mag}}$ magnitudes, independently for each winter. The sky background and its correlations with other parameters are shown in Fig. 9. The first row shows the sky background as a function of time. The Sun increases the sky background every day around noon and the
Moon cycle is clearly visible, as already reported in Crouzet et al. (2010). The second row displays the histograms of sky values. They are composed of three regions: a broad distribution centered around the nominal dark sky value (∼20.0 mag arcsec⁻²), a region between 19.5 and 18 mag arcsec⁻² with an increased number of counts that correspond to the illumination around full Moon, and a long tail for values brighter than 18 mag arcsec⁻² that correspond to the illumination by the Sun around noon. To measure the average dark sky background at Dome C, we combined the distributions of the four winters and fitted the main peak with a Gaussian function. We did this calculation in flux units and converted the mean and 1σ values back to magnitudes. We find a dark sky background at Dome C of 20.0±0.3 mag arcsec⁻² in the 579–642 nm bandpass. The uncertainty is at least partly due to real variations of the dark sky background: in 2011, the sky is darker around mid-winter (∼20.5 mag arcsec⁻²) and brighter closer to the equinoxes (∼19.5 mag arcsec⁻²). This trend is less clear for the other winters due to less consistent and less complete coverage. For comparison, the dark sky background in mag arcsec⁻² in the R band is 20.4 at Mauna Kea⁵, 20.9 at Paranal (Patat 2008), 21.0 at La Palma⁶, and 21.2 at Cerro Tololo and La Silla (Kriscuianas et al. 2007). Thus, the sky background at Dome C is comparable to that of the best observing sites.

The sky background as a function of the Sun altitude $h_{\text{Sun}}$ is shown in the third row of Fig. 9. A limit $h_{\text{Sun}} < -15^\circ$ can be used to define dark sky conditions, whereas the sky background increases quickly for $h_{\text{Sun}} > -13^\circ$. Finally, we calculated the lightcurve dispersion at each epoch $\sigma_e$, using the stars in the magnitude interval 12–13.5, which represents the photometric quality at that epoch (see Sect. 5). We show $\sigma_e$ as a function of the sky background and Sun altitude in the fourth and fifth rows, respectively. We find that $\sigma_e$ starts increasing noticeably when the sky background becomes brighter than 18 mag arcsec⁻². We find that $\sigma_e$ remains at its average level while $h_{\text{Sun}} < -12^\circ$ and increases steeply for $h_{\text{Sun}} > -10^\circ$. These numbers give the conditions under which good quality photometry can be achieved at Dome C as seen by our 10 cm diameter instrument.

5. Photometric quality of Dome C

The long, continuous, and consistent database provided by ASTEP South is well suited to measure the weather quality during the winter at Dome C. The clear sky fraction for the winter 2008 was measured by Crouzet et al. (2010) using the number of detected stars in the field of view. Here, we use the quality of the lightcurves themselves to derive the photometric quality. In the following calculations, we used only the subset of stars with R magnitudes between 12 and 13.5 (1090 control stars). The final lightcurves are in the form of magnitude residuals. We calculated the standard deviation of the residuals at each epoch $\sigma_e$ using the control stars. This parameter provides a direct measurement of the data quality at each epoch: large dispersions in the residuals correspond to low quality epochs. As shown in Fig. 12, the $\sigma_e$ distribution is composed of a bulk and a tail containing the high and low quality data, respectively. To differentiate the bulk and the tail, we built a model for $\sigma_e$ and compared it to the measured distribution. In the following, $\sigma_{e,b}$ and $\sigma_{e,t}$ refer to the measured and theoretical $\sigma_e$, respectively. We calculated $\sigma_{e,t}$ in several steps as described next.

5.1. Theoretical noise

First, we calculated the theoretical noise that was expected for each star at each epoch taking into account the Poisson noise, read-out noise, and sky background noise. We did not consider scintillation noise because it is negligible in the magnitude range 12–13.5 (see Fig. 6). Each star had its own photometric aperture which remained fixed in size for the four years of data. The fraction of the PSF included in the aperture depended on its FWHM, which varied in time. We calculated this fraction for each star at each epoch by integrating numerically the PSF over the photometric aperture. The PSF was taken as a two-dimensional Gaussian function elongated in one direction, its FWHM was that measured by ASTEP South, and the elongation was calculated as a function of the distance of the star from the celestial south pole. The photometric aperture was the same as used in the reduction pipeline.

We calculated the nominal flux for each star as the median of the flux measured under excellent conditions, which we defined as non-flagged epochs falling within the 20% largest number of stars and the 20% smallest FWHM. Then, we calculated the theoretical flux for each star at each epoch by scaling the nominal flux to the fraction of PSF within the aperture, and used it to derive theoretical estimates of the stellar Poisson noise, read-out noise, and sky background noise, relatively to the stellar signal. In these calculations and in the following, we included a 10% uncertainty on the measured FWHM and sky background values to account for measurement errors.

As shown in Sect. 3.7, these theoretical noises do not explain fully the measured lightcurves’ RMS. Thus, we added unidentified noise sources and split their contributions into two components $\sigma_w$ and $\sigma_r$, defined as:

$$\sigma_w^2 = \sigma_p^2 - \sigma_r^2 \quad \text{and} \quad \sigma_r^2 = \sigma_e^2 - \sigma_p^2$$

where $\sigma_r$ is the theoretical RMS, $\sigma_p$ is the point-to-point RMS, and $\sigma_e$ is the true RMS. $\sigma_w$ corresponds to a white (uncor-
lated) noise component and \( \sigma_r \) as a red (correlated) noise component. We obtained \( \sigma_w \) and \( \sigma_r \) for each star directly from their lightcurve (see Fig. 6); these lightcurves were cleaned from bad data points including a cut-off in \( \sigma_e \) so they correspond to good weather conditions. We expressed the noise \( \sigma_{s,e} \) for star \( s \) at epoch \( e \) as:

\[
\sigma_{s,e}^2 = \sigma_{\text{phot},s,e}^2 + \sigma_{\text{read},s,e}^2 + \sigma_{\text{sky},s,e}^2 + \sigma_{w,s}^2 + \sigma_{r,s}^2
\]

where \( \sigma_{\text{phot},s,e} \), \( \sigma_{\text{read},s,e} \), and \( \sigma_{\text{sky},s,e} \) are the stellar photon noise, read-out noise, and sky background noise, and \( \sigma_{w,s} \) and \( \sigma_{r,s} \) are the additional white and red noise contributions.

Finally, we computed the theoretical noise for epoch \( e \), \( \sigma_{e,t} \), as the median of \( \sigma_{s,e} \) over the control stars. As shown in Fig. 10, the resulting \( \sigma_{e,t} \) distribution is well centered on the \( \sigma_{e,m} \) distribution but is very narrow. This is because the correlation of \( \sigma_{e,s} \) between stars (the correlative nature of \( \sigma_{r,s} \)) was not accounted for: \( \sigma_{r,s} \) is only an amplitude. Increasing the random errors on the FHWM and sky background measurements broadened only slightly the \( \sigma_{e,t} \) distribution. In fact, more complex correlated noise was present and was not captured yet.

### 5.2. Addition of a blind correlative component

A complete model of the variations affecting the ASTEP South lightcurves was beyond the scope of this work. Instead, we used an empirical approach and included a correlative component in the \( \sigma_{r} \) parameter. We modeled the correlations in the follow-
\[ R = \sigma_{e,m} \left(1 + N(\beta)\right) \]  

(5)

The red noise value \( \sigma_{R,s} \) for star \( s \) was taken as the realization of \( R \) on the segments and remained constant on each segment. The theoretical noise \( \sigma_{e,s} \) for star \( s \) at epoch \( e \) was then defined as:

\[ \sigma_{e,s}^2 = \sigma_{\text{phot},s,e}^2 + \sigma_{\text{read},s,e}^2 + \sigma_{\text{sky},s,e}^2 + \sigma_{w,s}^2 + \sigma_{R,s,e}^2 \]  

(6)

where \( \sigma_{R,s,e} \) was then correlated between stars. The theoretical noise for epoch \( e \), \( \sigma_{e,i} \), was taken as the median of \( \sigma_{e,s} \) over the control stars.

The amount of correlation encoded in the \( \beta \) parameter was unknown; the larger it was, the broader was the \( \sigma_{e,i} \) distribution. We adjusted \( \beta \) by matching the left side of the \( \sigma_{e,i} \) histogram to that of \( \sigma_{e,m} \). Indeed, the broadening on the left side of the \( \sigma_{e,m} \) histogram was due to noise that was present even in good weather conditions, whereas the broadening on the right side was also due to bad weather, high sky background, etc. The resulting \( \sigma_{e,i} \) distribution reproduced well the bulk of the \( \sigma_{e,m} \) distribution, as shown in Fig. 10.

In this model, the sizes of the segments did not matter as long as the stars were affected in the same way. The \( \beta \) parameter represented the amount of correlated noise at a given time in the star domain, whereas the correlated noise is generally measured on a lightcurve, that is for a given star in the time domain (e.g., Gillon et al. 2006). The value of \( \beta \) depended on the number of correlated stars: if less stars were correlated, \( \beta \) had to be larger to reproduce the \( \sigma_{e,m} \) distribution; different groups of stars may also have had different correlations. Algorithms such as SYSREM (Tamuz et al. 2005), trend filtering algorithm (TFA, Kovács et al. 2005) or others using a principal component analysis aim at identifying and removing such correlations (although the use of SYSREM on the ASTEP South data did not improve the lightcurves compared to our own procedures). Our approach was equivalent to convoluting the \( \sigma_{e,i} \) distribution from Eq. 4 with a Gaussian function, but related this broadening to an amount of correlated noise that could be included in the noise budget. A more general approach based on the covariance matrix could also have been used (Pont et al. 2006).

### 5.3. Correlations with the FWHM and sky background

The tail of the \( \sigma_{e,m} \) histogram contained the bad weather data, but also included other noise sources which effects were not symmetrical with respect to the mode of the histogram and which were not included in our model. We investigated potential correlations of \( \sigma_{e,m} \) with the various recorded parameters for data taken under good weather conditions, and found clear correlations with the PSF FWHM and the sky background.

In spite of the correction by reference stars, FWHM variations affected the photometry more than predicted in Sect. 5.1. To derive the correlation between \( \sigma_{e,m} \) and the FWHM, we divided the FWHM measurements into ten intervals, and calculated the median FWHM and the median \( \sigma_{e,m} \) in each interval. In this process, we removed the \( \sigma_{e,m} \) outliers using a ten iteration sigma-clipping procedure with a threshold of 1.5 times the standard deviation, to ensure that the correlation was extracted from good weather data only. The extracted correlation function was denoted \( f \) and is shown for the winters 2010 and 2011 in Fig. 11. We interpolated \( f \) linearly to each measured FWHM (after including a 10% uncertainty on the measurements) and calculated the associated noise \( \sigma_{F,e} \) by:

\[ \sigma_{F,e}^2 = f^2 - f_0^2 \]  

(7)

where \( f_0 \) was the minimum of \( f \). This ensured that \( \sigma_{F,e} \) only accounted for additional noise caused by FWHM variations. We used the same \( \sigma_{F,e} \) value for all stars at a given epoch.

We used the same method for the sky background noise. We calculated the correlation function \( g \) using 22 intervals, with smaller sizes at lower sky values. \( g \) is shown for the winters 2010 and 2011 in Fig. 11. We interpolated \( g \) at each measured sky value (after including a 10% uncertainty on the measurements) and calculated the associated noise \( \sigma_{S,e} \) by:

\[ \sigma_{S,e}^2 = g^2 - g_0^2 \]  

(8)

where \( g_0 \) was the minimum of \( g \). This empirical model was close to the theoretical estimate from Poisson noise but with some differences at low sky values, possibly related to image calibration.

We also had to modify the theoretical relative photon noise, read-out noise, and sky noise terms in Equations 4 and 6 because part of their contributions was then contained in \( \sigma_{F,e} \) and \( \sigma_{S,e} \). To account for this, we calculated the stellar fluxes for a unique photometric aperture, but they were equal for all epochs. The FWHM and sky background variations were also responsible for a certain amount of red noise, so the contribution of \( \sigma_{R,s,e} \) had to be reduced. Some amount may have been considered as white noise (larger Poisson noise for larger FWHM or sky background), but we reduced only \( \sigma_{R,s,e} \) and not \( \sigma_{e} \) (the final result was equivalent). We defined a new red noise distribution \( R' \) as:

\[ R' = \alpha R \]  

(9)
as illustrated in Fig. 12. We repeated this procedure 100 times with different random distributions for the definition of the correlation segments, for R, and for the uncertainties on the FWHM and sky background measurements. We took the mean and standard deviation of the results to obtain the photometric and veiled weather statistics and their uncertainties. In addition, small fractions of the lightcurves were of reasonable quality but were not well corrected by the initial one-sidereal day period variation removal. They were kept but flagged early in the pipeline, and did not have a $\sigma_e$ value. Whether these data were affected by instrumental effects or poor weather conditions could not be inferred by our method. We classified these periods manually by interpreting the variations using the recorded parameters, the daily reports of the winter-over crew, and knowledge of the instrument behavior. We adopted a relative uncertainty of 10% for the corresponding fractions. These periods tend to increase the veiled weather fraction, and some of them correspond to bad weather. Furthermore, bad data were also eliminated during the photometric calibration process, and the instrument was stopped during “white-outs” (winter storms lasting one to a few days). We included these periods and adopted a relative uncertainty of 10% for the corresponding fractions. Finally, in the first part of 2009 (JD < 2455032), the instrument and lightcurves suffered extreme variations and our model did not reproduce the measured distribution accurately, so we derived the weather statistics only from the second part of the winter.

After these adjustments, we computed the final weather statistics for the winters at Dome C, which are reported in Table 5. The cumulative distributions of photometric weather, veiled weather, and bad weather plus white-out fractions are shown in Fig. 13. We also calculated the average weather statistics for the winter at Dome C by taking the mean and standard deviation of the results to obtain the photometric and veiled weather statistics and their uncertainties. In addition, small fractions of the lightcurves were of reasonable quality but were not well corrected by the initial one-sidereal day period variation removal. They were kept but flagged early in the pipeline, and did not have a $\sigma_e$ value. Whether these data were affected by instrumental effects or poor weather conditions could not be inferred by our method. We classified these periods manually by interpreting the variations using the recorded parameters, the daily reports of the winter-over crew, and knowledge of the instrument behavior. We adopted a relative uncertainty of 10% for the corresponding fractions. These periods tend to increase the veiled weather fraction, and some of them correspond to bad weather. Furthermore, bad data were also eliminated during the photometric calibration process, and the instrument was stopped during “white-outs” (winter storms lasting one to a few days). We included these periods and adopted a relative uncertainty of 10% for the corresponding fractions. Finally, in the first part of 2009 (JD < 2455032), the instrument and lightcurves suffered extreme variations and our model did not reproduce the measured distribution accurately, so we derived the weather statistics only from the second part of the winter.

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### 5.5. Validation

We validated the results of our method by verifying the epoch quality with respect to several parameters. To infer the quality of individual epochs, we calculated the ratio $\sigma_{e,m}/\sigma_{e,t}$ and we considered that an epoch is photometric or veiled if this ratio was below or above a certain limit, respectively. This limit was calculated for each winter to match the weather statistics. Fig.

#### Table 4: Model parameters $\beta$ and $\alpha$ for the epoch noise distribution $\sigma_{e,t}$ and ratio of the red noise contributions $r$. The values and uncertainties were obtained after repeating the procedure 100 times.

| Winter | $\beta$        | $\alpha$      | $r$  |
|--------|----------------|---------------|------|
| 2008   | $0.22 \pm 0.09$| $0.50 \pm 0.01$| 1.73 |
| 2009   | $0.17 \pm 0.07$| $0.38 \pm 0.01$| 2.44 |
| 2010   | $0.39 \pm 0.09$| $0.37 \pm 0.01$| 2.49 |
| 2011   | $0.21 \pm 0.02$| $0.65 \pm 0.01$| 1.18 |

#### 5.4. Weather statistics

We attributed to veiled weather the values of $\sigma_{e,m}$ in excess of our model. We defined the “photometric” and “veiled” weather fractions as the areas of $\sigma_{e,m}$ below and above $\sigma_{e,t}$ respectively,

$$\sigma^2_{e,c} = \sigma^2_{\text{phot},s} + \sigma^2_{\text{read},s} + \sigma^2_{\text{sky},s} + \sigma^2_{\text{psf},s} + \sigma^2_{\text{e},s} + \sigma^2_{\text{r},s,}\text{e}$$

As before, the theoretical noise for epoch $e$, $\sigma_{e,t}$, was taken as the median of $\sigma_{e,c}$ over the control stars. We derived $\alpha$ by matching the left sides of the $\sigma_{e,t}$ and $\sigma_{e,m}$ histograms. The final $\sigma_{e,t}$ distributions for each winter are shown in Fig. 12. We also defined the ratio $r$ of the joint contribution of $\sigma_{e,c}$ and $\sigma_{e,m}$ in the red noise budget compared to the unknown sources $\sigma_{e,c}$:

$$r = \sqrt{1/\alpha^2 - 1}$$

The two parameters of our model, $\beta$ and $\alpha$, and the ratio $r$ are reported for each winter in Table 4. For the four winters, $\sigma_{e,c}$ and $\sigma_{e,m}$ have a larger contribution than $\sigma_{e,s,}\text{r}$ ($r > 1$), so they are the main sources of correlated noise. For 2011, $r$ is lower than for the other winters, which is consistent with the stable FWHM during that winter. Following this approach, we could in principle have split entirely the red noise into its various components. However, no obvious correlations were found between $\sigma_{e,m}$ and other recorded parameters that were independent of the FWHM and the sky background, so we kept $\sigma_{e,\text{r,s,}}\text{e}$ as a global representation of the remaining components.

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As before, the theoretical noise for epoch $e$, $\sigma_{e,t}$, was taken as the median of $\sigma_{e,c}$ over the control stars. We derived $\alpha$ by matching the left sides of the $\sigma_{e,t}$ and $\sigma_{e,m}$ histograms. The final $\sigma_{e,t}$ distributions for each winter are shown in Fig. 12. We also defined the ratio $r$ of the joint contribution of $\sigma_{e,c}$ and $\sigma_{e,m}$ in the red noise budget compared to the unknown sources $\sigma_{e,c}$:

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Table 5: Weather statistics for the winter at Dome C as observed with ASTEP South. All numbers are in percent.

|                  | 2008      | 2009      | 2010      | 2011      | Average |
|------------------|-----------|-----------|-----------|-----------|---------|
| Excluding white-outs: |          |           |           |           |         |
| Photometric      | 75.6 ± 2.2 | 69.6 ± 1.2 | 75.3 ± 2.7 | 70.9 ± 1.2 | 72.9 ± 3.1 |
| Veiled           | 23.1 ± 2.2 | 25.8 ± 1.1 | 22.3 ± 1.7 | 23.3 ± 1.1 | 23.7 ± 1.5 |
| Bad              | 1.3 ± 0.1  | 4.5 ± 0.5  | 2.3 ± 0.2  | 5.8 ± 0.6  | 3.5 ± 2.0  |
| Including white-outs: |          |           |           |           |         |
| Photometric      | 63.0 ± 2.2 | 64.7 ± 1.2 | 72.5 ± 2.6 | 68.0 ± 1.2 | 67.1 ± 4.2 |
| Veiled           | 19.2 ± 1.8 | 24.0 ± 1.1 | 21.5 ± 1.6 | 22.4 ± 1.1 | 21.8 ± 2.0 |
| Bad              | 1.1 ± 0.1  | 4.2 ± 0.4  | 2.2 ± 0.2  | 5.5 ± 0.6  | 3.3 ± 2.0  |
| White-out        | 16.6 ± 1.7 | 7.1 ± 0.7  | 3.8 ± 0.4  | 4.1 ± 0.4  | 7.9 ± 6.0  |

Fig. 12: Normalized distribution of \( \sigma_e \) for the winters 2008, 2009, 2010, and 2011 (black lines). Periods of photometric and veiled weather as defined by our model are indicated in red and orange respectively.

Fig. 13: Cumulative normalized distribution of \( \sigma_e \) for the winters 2008, 2009, 2010, and 2011 (black lines). Periods of photometric and veiled weather are indicated in red and orange respectively, and include adjustments for periods without a \( \sigma_e \) value (see text). Periods of bad weather and white-out are indicated in white.

6. Duty cycle

The duty cycle of ASTEP South for the winters 2008 to 2011 is represented in Fig. 16: the limit due to the Sun elevation, the fraction of time where ASTEP South was operating, and the photometric, veiled, and bad weather fractions including white-outs are shown for each day of the winter. In total, we acquired 11415.1 hours of data with ASTEP South including 6699.8 hours under photometric conditions and nominal functioning of the instrument. Considering only times with a Sun elevation lower than \(-9^\circ\), these numbers are 9451.2 hours and 6308.5 hours respectively. The detail for each year is reported in Table 6. The increase of the operating time from the 2008 to the 2011 winter is evident and reflects the improvements of the system (hardware and software) made over the years. The first half of the 2008 winter was hampered by numerous technical issues while setting up the instrument. A long white out period occurred at the end of August and lasted for about ten days. In 2009, the be-
Crouzet et al.: ASTEP South

Fig. 14: Correlations of $\sigma_e$ with the PSF FWHM (top) and the sky background (middle and bottom) for the winters 2010 (left) and 2011 (right). Photometric and veiled epochs as inferred from our model are indicated in red and orange, respectively. Bad epochs are indicated in blue. The sky background plots are shown in two different scales.

Fig. 15: Correlations of the number of detected stars with $\sigma_e$ (top left), the FWHM (top right), and the sky background (bottom left and right, in two different scales) for the 2011 winter. Photometric and veiled epochs as inferred from our model are indicated in red and orange, respectively. Bad epochs are indicated in blue; most of them do not have a calculated $\sigma_e$.

Ginning of the winter was dedicated to setting up the instrument and improving the thermalization; data acquired during this time were not used in the lightcurves. In 2010, we lost two weeks of data taken in June, either in the data transfer between computers and external disks or while shipping the disks from Concordia to Nice. In 2011, the installation was fast and the system acquired high quality data nearly continuously for the whole winter, with almost no human interventions. This demonstrates that a simple instrument can operate almost continuously for several winters at Dome C. In addition, as noted in Crouzet et al. (2010), periods of bad weather are generally concentrated in white-out episodes lasting from one to a few days; between these episodes, high quality photometric observations can be achieved for extended periods of time. Overall, the weather statistics and duty cycle at Dome C demonstrate the high quality of this site for continuous photometric observations during the winter.

7. Study of variable stars and eclipsing binaries

7.1. Search for periodic signals

We searched for periodic signals in the lightcurves using the box least square (BLS) algorithm (Kovács et al. 2002). This program searches for square shapes representing transit-like events, and also detects stellar variations. We performed the search over a wide period range ($0.4 < P < 100$ days) which we divided into 15 intervals with sizes following approximately a logarithmic scale. We built a frequency comb for each interval using a constant spacing in $\Delta f/ f$, a number of frequencies set by the expected duty cycle, and we ran BLS in each interval separately. During the search, we applied locally a refined comb with nine times more frequencies, first around the forty best frequencies, then around the best frequency. This adaptive comb yielded a more precise determination of the best frequency. We ran BLS twice. After the first run, the histogram of periods showed peaks that were centered around integer or half-integer values, indicating harmonics of the “day-night” cycle (the sky background increases around noon even during the winter, see Sect. 4.3). Another peak around 29 days was due to the Moon cycle. We excluded periods around these peaks for the second BLS run. Finally, we combined the 15 intervals into three period ranges: $0.4 – 1$ day, $1 – 10$ days, and $10 – 100$ days. For each lightcurve, we compared the results obtained for the intervals within each range and kept only the period that yielded the largest power in the BLS spectrum. Every lightcurve ended up with a best period, but most of them did not contain a valid signal. We ranked them by decreasing signal to noise ratios, where the signal was calculated from the amplitude and duration of the variations and the noise was obtained from the standard deviation of the lightcurve residuals. Then, we inspected the lightcurves visually to identify variable stars, eclipsing binaries, and transiting exoplanet candidates. Examples of lightcurves of variable objects are shown in Fig. A.1.

7.2. Analysis of variable objects

We detected periodic signals in the lightcurves of 211 stars. A similar search was performed by the Chinese Small Telescope ARray (CSTAR, Yuan et al. 2008; Zhou et al. 2010) from Dome A, Antarctica, using data acquired during the 2008, 2009, and 2010 winters (Wang et al. 2011, 2013; Oelkers et al. 2015; Yang et al. 2015; Wang et al. 2015). We cross-matched our objects with the variable stars reported by CSTAR and by the GCVS (Samus’ et al. 2017), VSX (Watson et al. 2006), and ASAS (Pojmanski 2002; Richards et al. 2012) catalogs. The results are given in Table 7. We identified 96 objects that were not reported and we analyzed their lightcurves with the Period04 software program (Lenz & Breger 2005). First, we analyzed the 2011 data alone because they were more abundant and of
better quality, then we analyzed the four winters together. In the Fourier spectra, we found systematic peaks around 1.003 and 0.034 $d^{-1}$ caused by the day-night and Moon cycles, respectively (these frequencies were not completely avoided in our BLS search probably due to the frequency refinement procedure). We eliminated stars showing variations only at these frequencies. We also discarded stars with frequencies that were common to many stars, indicating non-astrophysical origin, and stars with signal to noise ratios that were too low to extract accurate periods and amplitudes. This left 42 objects with confirmed variability. We analyzed them year by year to distinguish variable stars and eclipsing binaries: the amplitude of variations due to stellar activity varies from year to year but remains the same for eclipsing binaries. We also analyzed the lightcurves of five known objects: we investigated in detail $\sigma$ Oct and HD 177482, we improved the parameters of EN Oct, and we re-classified two CSTAR variables as eclipsing binaries.

In this study, we computed the frequency and amplitude uncertainties using the formulae proposed by Montgomery & Odonoghue (1999) providing 3$\sigma$ estimates. These uncertainties correspond to ideal cases (constant noise, regular spacing, no gaps) and underestimate the true uncertainties. Based on our experience, we used the 3$\sigma$ values for our uncertainties.

Finally, we counted the number of previously reported variables in the ASTEP South field of view that were missed by our search (that were not in our list of 211 potential variables). We missed 103 objects including 90 from CSTAR and 13 from GCVS, VSX, or ASAS. Some were not in our initial target list, some were too faint, and some were simply not identified as variable. CSTAR is composed of four Schmidt-Cassegrain telescopes, each with a pupil entrance aperture of 145 mm and a field of view of 4.5$^\circ$ in diameter (20$''$ in total) equipped with different filters (g, r, i, and no filter). The CSTAR target list contains around 18,000 objects. Although technical issues generally prevented the acquisition of data by the four telescopes simultaneously, each CSTAR telescope already collects about twice more flux than ASTEP South. This set up can explain that many variables were detected by CSTAR and missed by ASTEP South. The photometric data reduction pipelines are also different, and Oelkers et al. (2015) use differential images. In addition, our search for periodic signals was primarily oriented toward the detection of transiting exoplanet candidates and we used only the BLS algorithm to search for variability. Other algorithms or metrics that were used for CSTAR such as the Welch-Stetson variability index (Stetson 1996), the analysis of variance method (Schwarzenberg-Czerny 1996), or the Lomb-Scargle method (Lomb 1976; Scargle 1982) may reveal more variables in the ASTEP South lightcurves. Besides, we restrained the search range to periods larger than 0.4 days, and most $\delta$ Scuti were not identified by their main frequency: some were found from their harmonics and some may have been missed. In contrast, the longer coverage by ASTEP South compared to CSTAR may explain that some variables were detected only by ASTEP South.

Our goal is not to detect all the variables in the ASTEP South field of view, which has already been studied extensively by CSTAR. Instead, we complement the CSTAR search and focus in the following on a few interesting objects that highlight the long photometric coverage that we achieved with ASTEP South.

7.2.1. $\sigma$ Oct

$\sigma$ Oct, also known as HR 7228 and HD 177482, is the brightest star in the field (V = 5.42). It is a $\delta$ Scuti reported first by McNally & Austin (1978) and re-observed only once by Coates et al. (1981). Its properties are poorly known as it had been observed...
Table 7: Summary of our search for variable objects in the ASTEP South lightcurves.

| Category                        | Value  |
|---------------------------------|--------|
| Number of target stars          | 5954   |
| Objects with detected variability| 211    |
| New objects                     | 96     |
| Confirmed variability           | 42     |
| New variable stars              | 34     |
| New eclipsing binaries          | 8      |
| Unconfirmed variability         | 54     |
| Known objects                   | 115    |
| From CSTAR                      | 95     |
| From GCVS, VSX, and ASAS        | 20     |
| Objects missed                  | 103    |
| From CSTAR                      | 90     |
| From GCVS, VSX, and ASAS        | 13     |

only four nights in total before the ASTEP South and CSTAR experiments. Its known pulsation period is on the order of 0.1 day with an amplitude of 15 mmag. The ASTEP South lightcurve of $\sigma$ Oct is shown in Fig. 17. We analyzed the data with the Period04 software program and detected 21 frequencies with signal to noise ratios larger than four (Fig. 18, Table B.1). Seventeen of them are between 8.023 and 11.756 d$^{-1}$. The dominant peak is at 10.493 d$^{-1}$ and corresponds to $\delta$ Scuti-type oscillations with an average amplitude of 3.73 mmag. Interestingly, four other frequencies are between 0.612 and 2.848 d$^{-1}$ with amplitudes below 0.7 mmag, and could correspond to $\gamma$ Dor-type pulsations. Thus, $\sigma$ Oct is apparently an hybrid $\delta$ Scuti – $\gamma$ Dor with very low amplitudes for the $\gamma$ Dor modes. We analyzed the ten main $\delta$ Scuti frequencies year by year (Fig. 19, Table B.2). Four of them show amplitude variations. In particular, the amplitude of the dominant frequency decreases from 9.74 to 2.76 mmag (a factor of 3.5) between 2008 and 2011, and this decrease is not compensated by an equivalent increase of the amplitudes of other frequencies. Thus, the cause is not an energy transfer between oscillation modes but is an actual diminution of the total amplitude of the star’s oscillations. A similar behavior was found in 603 $\delta$ Scuti out of 983 using four year lightcurves from the Kepler mission and no explanation is known to date (Bowman et al. 2016).

7.2.2. HD 184465

HD 184465 is a star of $V$ magnitude 8.82 and spectral type A1 V, located in the $\delta$ Scuti instability strip. Its ASTEP South lightcurve and Fourier spectrum are shown in Figs. 20 and 21. We detected ten $\delta$ Scuti-type frequencies with signal to noise ratios greater than four and we analyzed them year by year (Fig. 22, Table B.3). The 9.671 d$^{-1}$ frequency largely dominates the Fourier spectrum, and its amplitude increases from 0.9 to 6.02 mmag (a factor of 6.7) between 2008 and 2011. This increase is not compensated by an equivalent decrease of the amplitudes of other frequencies. The amplitudes of other frequencies are generally too small to detect significant variations between years. We analyzed the 2011 data by 30-day segments (Table B.4, Fig. 22). The amplitudes of the 9.671 d$^{-1}$ and 6.465 d$^{-1}$ frequencies vary at similar rates in 2011 compared to the four winters together. Interestingly, the amplitude of the 8.691 d$^{-1}$ frequency increases in 2011 from 0.7 to 2.7 mmag (a factor of 3.8). Overall, this star does not show energy transfer between os-

Fig. 17: Lightcurve of $\sigma$ Oct as observed with ASTEP South over four winters, phase-folded at a period of 0.0953 day. The black points correspond to a time sampling of ten minutes. The red line is a binning with 100 points over the phase range. The period and $V$ magnitude are indicated on the plot.

Fig. 18: Fourier spectrum of $\sigma$ Oct obtained from its four winter lightcurve.

Fig. 19: Amplitude variations of the ten main $\delta$ Scuti frequencies of $\sigma$ Oct from 2008 to 2011, with one point per winter. The $3\sigma$ uncertainties are shown for the main frequency and are the same for all frequencies.
cillations modes, but shows actual variations of the amplitude of its oscillations on several timescales.

![Graph showing normalized flux with phase and period](image)

Fig. 20: Lightcurve of HD 184465 as observed with ASTEP South over four winters, phase-folded at a period of 0.1034 day. The black points correspond to a time sampling of ten minutes. The red line is a binning with 100 points over the phase range. The period and V magnitude are indicated on the plot.

![Fourier spectrum of HD 184465](image)

Fig. 21: Fourier spectrum of HD 184465 obtained from its four winter lightcurve.

### 7.2.3. New variable stars and eclipsing binaries

We identified 34 new variable stars (Table C.1). Their amplitudes are typically between 1 and 5 mmag and several have periods longer than ten days. More than half are active stars with variations induced by their rotation. In many cases, the signal to noise ratio is too low to identify the precise type of variable. We also identified eight new eclipsing binaries and reclassified two CSTAR variables as eclipsing binaries (Table C.2). Six of them have orbital periods between 23 and 81 days, which are much longer than those usually detected from the ground. In addition, we refined the period and eclipse depth of the known eclipsing binary EN Oct compared to previous studies (Knigge & Bauernfeind 1967; Otero 2004). We also detected variable stars and eclipsing binaries already reported by CSTAR and other catalogs, so we do not report them here. Figure A.1 shows examples of lightcurves of variable stars and eclipsing binaries obtained with ASTEP South. Despite challenges in the photometric reduction due to the unusual observing mode of ASTEP South, these lightcurves are of good quality, and the large amount of data enables the detection of variations over a wide range of magnitudes and periods (in Fig. A.1, the V magnitudes range from 6.52 to 15.04 and the periods range from 0.12 to 40.53 days). Overall, these detections demonstrate the benefit of the long photometric coverage that can be achieved during the winter at Dome C.

### 7.2.4. Inventory of variable objects

We gathered the new variable stars and eclipsing binaries detected by ASTEP South with those reported by CSTAR, GCVS, VSX, and ASAS to make an inventory of variable objects in the ASTEP South field of view. Their classification is given in Table 8. Further details on these objects can be found in Wang et al. (2011, 2013); Oelkers et al. (2015); Yang et al. (2015); Wang et al. (2015); Samus’ et al. (2017); Watson et al. (2006); Pojmanski (2002); Richards et al. (2012) and in Tables C.1 and C.2 of this paper.
Table 8: Inventory of variable objects in the ASTEP South field of view (Dec < −88°) including those detected by ASTEP South and those reported by CSTAR, GCVS, VSX, and ASAS.

| Type of Object        | Number |
|-----------------------|--------|
| Variable stars        | 209    |
| VRot                  | 96     |
| Ceph                  | 1      |
| Delta Scuti           | 19     |
| Gamma Dor             | 10     |
| RR Lyrae              | 4      |
| Mira                  | 2      |
| Multiperiodic         | 7      |
| Irregular             | 38     |
| Unclassified          | 32     |
| Eclipsing binaries    | 47     |
| EA                    | 19     |
| EB                    | 10     |
| EW                    | 18     |
| Transit-like          | 4      |

Notes. “VRot” are rotational variables and include the ACV, BY, SP, ELL classes and the “VRot?” of Table C.1. “Unclassified” include the “Var” of Table C.1. “Transit-like” objects are those reported by CSTAR; the transit candidates identified by ASTEP South are not reported here and are under investigation. This table does not include the 54 objects initially suspected as variable with ASTEP South that were not confirmed by a finer analysis (see Sect. 7.2).

7.2.5. A 75.6 day period G–M eclipsing binary

Eclipsing binaries are the most common by-products of transiting exoplanet searches. In our search for transiting objects, we detected an eclipsing binary with a 75.6 day orbital period and a 13.6 hour primary eclipse (star ID: UCAC3 2-3056, fmag: 12.76; see Table C.2). These period and duration are much longer than those usually detected from the ground, which shows the benefit of the exceptional phase coverage that we achieved from Dome C. This object was not detected by Period04 because of its too low signal to noise ratio but it was among the transit candidates that we identified from the BLS search and visual inspection. We conducted follow-up observations in radial velocity and photometry, which confirmed the orbital period and the presence of the eclipses. This system is a G9–M2.5 eclipsing binary with an eccentric orbit (e = 0.26) and a relatively large semi-major axis (a = 0.42 au). Thus, tidal coupling between both stars is weak and the M dwarf should behave as a single star. This object will provide an ideal benchmark to constrain the evolution models of low mass stars, and will be presented in a forthcoming paper (Crouzet et al. in prep).

8. Conclusion

We presented the full analysis of four winters of observations collected with ASTEP South, a 10 cm refractor installed at Dome C, Antarctica. The instrument ran continuously for most of the winters. The functioning of the instrument and the lightcurve quality improved over the years. A main limitation was shutter malfunctions that started after two winters and required specific temperature adjustments and image calibration methods. Another factor that affected the data quality was imperfect thermalisation, in particular during the first half of the 2009 winter, caused by deficient air circulation leading to variable temperature and gradients inside the box. This resulted in unstable PSFs and poor quality lightcurves for these periods. By acquiring knowledge of the instrument’s behavior over the years and improving the experiment (hardware and software), we were able to observe longer and longer, from half of the winter in 2008 to the full winter with almost no intervention in 2011. As previously shown for the 2008 winter, we confirmed that the instrument was very stable over the ice and that PSF stability was achieved when the thermalisation was nominal. We also measured the sky background magnitude and its variations over the winters.

We developed a specific data reduction pipeline to take into account the motion and elongated shape of the PSFs on the CCD. We built the lightcurves of nearly 6000 stars over four winters, which are available publicly. The experimental design was particularly sensitive to flat-field errors, which we mitigated by removing the one-sidereal day period variations in the lightcurves. Most of the data are of good quality. The final lightcurve RMS is standard for a 10 cm instrument, and the large quantity of data yields a drastic improvement of the precision when folding the lightcurves at given periods and binning the data points. We measured the photometric quality of Dome C from the lightcurves themselves using a method that we developed specifically for this work, which may be applied to other experiments.

We detected periodic signal in the lightcurves of around two hundred objects. Half of them were already reported by the CSTAR experiment. We discovered new variable stars and eclipsing binaries, and conducted a detailed study of a few objects that was enabled by the long and continuous coverage of our observations. These results demonstrate the high quality of Dome C for photometry in the visible and the technical feasibility of running astronomical instruments for several years at this site.

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Article number, page 18 of 23page.23
Fig. A.1: Examples of phase-folded lightcurves of variable stars and eclipsing binaries obtained with ASTEP South, sorted by magnitude. The black points correspond to a time sampling of ten minutes. A binning with 100 points over the phase range is shown as a red line, except for eclipsing binaries with narrow eclipses. The star ID, the period in days, and the V magnitude are indicated.
Appendix B: Analysis of δ Scuti pulsations of σ Oct and HD 184465

Table B.1: Twenty-one main frequencies detected in the Fourier spectrum of σ Oct with their amplitudes and phases. The reference for the phase is BJD 2454600. Numbers in parentheses indicate the 3σ uncertainties on the last two digits.

| Frequency [d\(^{-1}\)] | Amplitude [mmag] | Phase [rad] |
|-------------------------|-----------------|-------------|
| 10.492693 (13)          | 3.73 (11)       | 0.449 (30)  |
| 9.719016 (16)           | 3.11 (11)       | 2.972 (36)  |
| 10.741796 (18)          | 2.77 (11)       | 2.953 (40)  |
| 8.799398 (19)           | 2.61 (11)       | 4.475 (42)  |
| 10.251893 (20)          | 2.47 (11)       | 4.517 (45)  |
| 11.429052 (22)          | 2.27 (11)       | 2.818 (49)  |
| 9.361590 (28)           | 1.80 (11)       | 2.375 (61)  |
| 9.771399 (30)           | 1.65 (11)       | 1.183 (67)  |
| 10.440784 (37)          | 1.37 (11)       | 6.276 (81)  |
| 10.058734 (52)          | 0.96 (11)       | 1.24 (12)   |
| 11.756385 (69)          | 0.73 (11)       | 4.06 (15)   |
| 9.139954 (75)           | 0.67 (11)       | 1.88 (17)   |
| 0.960644 (81)           | 0.62 (11)       | 3.39 (18)   |
| 0.611525 (82)           | 0.61 (11)       | 3.31 (18)   |
| 0.898446 (85)           | 0.59 (11)       | 4.89 (19)   |
| 11.189760 (93)          | 0.54 (11)       | 5.06 (21)   |
| 2.848108 (99)           | 0.50 (11)       | 2.09 (22)   |
| 11.26188 (10)           | 0.49 (11)       | 4.08 (23)   |
| 8.23509 (11)            | 0.47 (11)       | 5.52 (24)   |
| 8.02296 (14)            | 0.35 (11)       | 2.15 (32)   |
| 11.02859 (16)           | 0.32 (11)       | 5.41 (35)   |

Table B.2: Amplitude of the ten main δ Scuti frequencies found in the Fourier spectrum of σ Oct for each winter and for the four winters together. Numbers in parentheses indicate the 3σ uncertainties on the last two digits. The amplitudes are in mmag.

| Frequency [d\(^{-1}\)] | 2008     | 2009     | 2010     | 2011     | All      |
|-------------------------|----------|----------|----------|----------|----------|
| 10.492693 (13)          | 9.69 (32)| 5.42 (26)| 3.01 (22)| 0.99 (16)| 3.73 (11)|
| 9.719016 (16)           | 2.58 (32)| 2.78 (26)| 3.07 (22)| 3.32 (16)| 3.11 (11)|
| 10.741796 (18)          | 2.84 (32)| 2.74 (26)| 2.67 (22)| 2.79 (16)| 2.77 (11)|
| 8.799398 (19)           | 2.91 (32)| 2.56 (26)| 2.56 (22)| 2.52 (16)| 2.61 (11)|
| 10.251893 (20)          | 2.29 (32)| 2.14 (26)| 2.27 (22)| 2.58 (16)| 2.47 (11)|
| 11.429052 (22)          | 2.39 (32)| 2.24 (26)| 2.12 (22)| 2.53 (16)| 2.27 (11)|
| 9.361590 (28)           | 2.99 (32)| 1.67 (26)| 1.51 (22)| 1.55 (16)| 1.80 (11)|
| 9.771399 (30)           | 1.56 (32)| 1.55 (26)| 1.82 (22)| 1.75 (16)| 1.65 (11)|
| 10.440784 (37)          | 1.21 (32)| 1.24 (26)| 1.36 (22)| 1.72 (16)| 1.37 (11)|
| 10.058734 (52)          | 1.38 (32)| 1.31 (26)| 0.93 (22)| 0.77 (16)| 0.96 (11)|
Table B.3: Amplitude of the ten $\delta$ Scuti frequencies with signal to noise ratios greater than four found in the Fourier spectrum of HD 184465 for each winter and for the four winters together. Numbers in parentheses indicate the 3$\sigma$ uncertainties on the last two digits. The amplitudes are in mmag.

| Frequency [d$^{-1}$] | 2008 | 2009 | 2010 | 2011 | All |
|----------------------|------|------|------|------|-----|
| 9.671458 (12)        | 0.86 (32) | 2.24 (24) | 3.60 (19) | 5.98 (14) | 3.94 (10) |
| 6.465094 (56)        | 1.33 (32) | 0.90 (24) | 0.93 (19) | 0.67 (14) | 0.85 (10) |
| 8.691272 (64)        | 0.26 (32) | 0.34 (24) | 0.28 (19) | 1.41 (14) | 0.74 (10) |
| 4.934047 (69)        | 0.79 (32) | 0.57 (24) | 0.72 (19) | 0.69 (14) | 0.68 (10) |
| 3.785687 (87)        | 0.36 (32) | 0.38 (24) | 0.63 (19) | 0.61 (14) | 0.55 (10) |
| 14.20371 (12)        | 0.07 (32) | 0.06 (24) | 0.19 (19) | 0.74 (14) | 0.40 (10) |
| 4.35554 (13)         | 0.24 (32) | 0.40 (24) | 0.36 (19) | 0.53 (14) | 0.37 (10) |
| 7.70589 (18)         | 0.26 (32) | 0.38 (24) | 0.38 (19) | 0.18 (14) | 0.26 (10) |
| 10.03363 (25)        | 0.39 (32) | 0.29 (24) | 0.19 (19) | 0.32 (14) | 0.19 (10) |
| 11.82126 (26)        | 0.19 (32) | 0.27 (24) | 0.10 (19) | 0.22 (14) | 0.19 (10) |

Table B.4: Amplitude variations of three $\delta$ Scuti frequencies of HD 184465 during the four winters, with two time intervals for each of the 2008, 2009, and 2010 winters and six 30-day intervals for the 2011 winter. Dates are the central time of each interval. Numbers in parentheses indicate the 3$\sigma$ uncertainties on the last two digits for the amplitudes (they are the same for all amplitudes at a given time interval). The amplitudes are in mmag.

| JD - 2450000 | Frequency [d$^{-1}$] |
|--------------|----------------------|
|              | 9.671 | 8.692 | 6.465 |
| 4652         | 0.84  | 0.32  | 1.30  | (36) |
| 4702         | 1.04  | 0.41  | 1.40  | (52) |
| 5003         | 2.23  | 0.39  | 0.79  | (35) |
| 5074         | 2.28  | 0.38  | 0.85  | (31) |
| 5329         | 3.27  | 0.35  | 0.81  | (32) |
| 5419         | 3.91  | 0.53  | 0.99  | (25) |
| 5652         | 5.59  | 0.76  | 0.75  | (39) |
| 5682         | 5.76  | 0.65  | 0.83  | (40) |
| 5712         | 5.59  | 1.37  | 0.77  | (26) |
| 5742         | 6.12  | 1.24  | 0.49  | (29) |
| 5772         | 6.02  | 1.88  | 0.89  | (33) |
| 5802         | 6.10  | 2.74  | 0.61  | (51) |
## Appendix C: New variables stars and eclipsing binaries identified with ASTEP South

Table C.1: New variable stars identified with ASTEP South. The ID, coordinates, and magnitudes $f_{\text{mag}}$ are extracted from the UCAC3 catalog. For each object, we report the frequency, the corresponding period, the semi-amplitude of the variations, and the type of variable. Numbers in parentheses indicate the $3\sigma$ uncertainties on the last digit for the frequencies, and the $3\sigma$ uncertainties for the amplitudes. “Var” indicates that the type of variable cannot be determined and “?” indicates a tentative classification.

| ID (UCAC3) | RA (J2000) | Dec (J2000) | $f_{\text{mag}}$ [mag] | Frequency [d$^{-1}$] | Period [d] | Amplitude [mmag] | Type |
|------------|------------|-------------|------------------------|---------------------|-----------|-----------------|------|
| 004-005139 | 13:31:52.604 | -88:22:17.75 | 10.23 | 0.14252 (3) | 7.02 | 2.3 (0.1) | Var |
| 003-002049 | 07:54:08.982 | -88:31:04.01 | 10.91 | 5.83351 (6) | 0.17 | 1.1 (0.1) | δ Scuti |
| 004-004378 | 15:31:11.257 | -88:31:17.40 | 10.93 | 0.10227 (3) | 9.78 | 3.7 (0.2) | V Rot? |
| 004-006257 | 16:13:29.743 | -88:12:25.00 | 11.82 | 0.08001 (2) | 12.50 | 4.5 (0.2) | Var |
| 002-000843 | 05:12:39.072 | -89:10:02.73 | 12.18 | 0.13586 (3) | 7.36 | 3.0 (0.5) | V Rot |
| 002-000214 | 01:15:41.760 | -89:00:19.72 | 12.21 | 0.02587 (2) | 38.65 | 6.1 (0.2) | V Rot |
| 002-000853 | 05:17:21.064 | -89:21:19.78 | 12.25 | 1.05540 (6) | 0.95 | 1.6 (0.2) | V Rot? |
| 004-000689 | 01:57:52.063 | -88:13:15.56 | 12.55 | 0.12161 (4) | 8.22 | 3.8 (0.3) | V Rot? |
| 003-002304 | 08:39:17.699 | -88:44:20.12 | 12.59 | 0.08060 (2) | 12.41 | 5.3 (0.3) | Var |
| 003-004226 | 15:05:46.232 | -88:39:23.33 | 12.67 | 0.02679 (3) | 37.33 | 3.9 (0.3) | V Rot |
| 003-005314 | 19:03:58.030 | -88:59:04.65 | 12.69 | 0.02492 (3) | 40.13 | 3.4 (0.2) | V Rot? |
| 001-000287 | 05:43:45.773 | -89:33:57.57 | 12.69 | 0.10116 (3) | 9.89 | 4.5 (0.2) | Var |
| 001-000176 | 03:35:19.894 | -89:41:08.84 | 12.80 | 0.04475 (3) | 22.35 | 4.0 (0.2) | Var |
| 003-001672 | 06:30:27.586 | -88:59:11.08 | 12.83 | 0.04459 (2) | 22.43 | 5.6 (0.2) | Var |
| 003-003813 | 13:42:59.863 | -88:53:36.27 | 12.96 | 0.07838 (2) | 12.76 | 6.7 (0.3) | V Rot |
| 003-000054 | 00:16:06.515 | -88:34:08.37 | 12.98 | 0.01286 (1) | 77.76 | 4.6 (0.4) | V Rot? |
| 002-005588 | 03:32:17.485 | -89:15:06.15 | 13.51 | 0.09741 (4) | 10.27 | 4.6 (0.4) | V Rot? |
| 003-001584 | 06:12:58.011 | -88:35:13.08 | 13.52 | 0.12314 (4) | 8.12 | 6.5 (0.5) | V Rot? |
| 003-000486 | 01:59:49.979 | -88:54:05.91 | 13.64 | 0.02878 (2) | 34.75 | 13.3 (0.6) | V Rot |
| 002-002489 | 14:58:37.642 | -89:07:37.09 | 13.81 | 0.04184 (2) | 23.90 | 18.5 (0.6) | V Rot |
| 003-004679 | 16:35:40.811 | -88:43:46.12 | 13.83 | 0.02809 (5) | 35.60 | 88.9 (1.0) | V Rot |
| 003-004666 | 16:32:18.171 | -88:36:33.72 | 14.21 | 0.02950 (3) | 33.90 | 18.4 (1.1) | V Rot |
| 004-004408 | 11:55:09.767 | -88:23:58.50 | 14.32 | 0.23194 (4) | 4.31 | 7.7 (0.7) | V Rot? |
| 004-004603 | 12:17:46.280 | -88:05:49.99 | 14.44 | 1.03020 (5) | 0.97 | 9.6 (1.1) | Var |
| 004-007265 | 19:04:41.581 | -88:29:17.88 | 14.46 | 2.11961 (4) | 0.47 | 5.4 (0.5) | Var |
| 003-003146 | 11:28:48.149 | -88:50:41.89 | 14.48 | 0.42148 (4) | 2.37 | 11.5 (1.0) | Var |
| 002-001119 | 06:54:02.979 | -89:05:10.41 | 14.54 | 5.89217 (4) | 0.17 | 10.8 (1.0) | Var |
| 003-006325 | 23:21:41.025 | -88:50:24.29 | 14.91 | 0.64068 (2) | 1.56 | 35.6 (1.4) | Var |
| 001-000349 | 06:48:03.486 | -89:32:05.52 | 14.99 | 5.64772 (6) | 0.18 | 97.0 (1.3) | δ Scuti |
| 001-000856 | 15:36:30.014 | -89:47:17.82 | 15.05 | 0.80258 (3) | 1.25 | 18.9 (1.2) | V Rot |
| 002-001114 | 06:52:18.556 | -89:01:33.13 | 15.09 | 0.12765 (2) | 7.83 | 14.9 (0.6) | Var |
| 002-003252 | 19:53:39.345 | -89:01:49.63 | 15.20 | 1.82823 (2) | 0.55 | 43.0 (1.4) | V Rot |
| 002-001915 | 11:32:03.119 | -89:04:57.12 | 15.21 | 0.95488 (2) | 1.05 | 28.6 (1.4) | Var |
| 002-001216 | 07:28:33.245 | -89:03:48.27 | 15.69 | 0.77189 (3) | 1.30 | 14.1 (0.9) | Var |
Table C.2: New eclipsing binaries identified with ASTEP South and known objects with improved classification and parameters. The ID, coordinates, and magnitudes $f_{\text{mag}}$ are extracted from the UCAC3 catalog. For each object, we report the period, the mid-eclipse date of the first detected eclipse, the depth of the primary eclipse (or of the secondary eclipse in the potential case of eccentric binaries with no primary eclipses), and the type of eclipsing binary. Numbers in parentheses indicate the 3σ uncertainties on the last two digits for the periods and ephemerides, and the 3σ uncertainties for the depths.

| ID (UCAC3) | RA (J2000) | Dec (J2000) | $f_{\text{mag}}$ [mag] | Period [d] | T$_0$ [BJD] | Depth [m mag] | Type |
|------------|------------|-------------|-------------------------|------------|-------------|---------------|------|
| 003-001765 | 06:53:22.343 | -88:34:04.28 | 10.04 | 0.4574110 (48) | 4627.4151 (34) | 2.6 (0.4) | EB |
| 002-002671 | 16:04:20.122 | -89:00:40.74 | 10.99 | 13.4892 (25) | 4627.70 (13) | 5.7 (0.4) | EB |
| 002-003763 | 23:23:51.240 | -89:25:17.62 | 11.25 | 23.3853 (53) | 4636.35 (18) | 8.6 (0.4) | EB |
| 004-006956 | 18:03:31.003 | -88:16:36.79 | 11.86 | 36.263 (24) | 4634.18 (23) | 5.9 (0.4) | EB |
| 004-000272 | 00:45:26.547 | -88:25:21.03 | 11.98 | 34.559 (19) | 4627.69 (19) | 15.0 (0.7) | EB |
| 001-000823 | 14:50:14.729 | -89:46:58.96 | 12.18 | 0.28891080 (20) | 4627.2346 (02) | 479 (8) | EW |
| 001-000701 | 12:41:48.571 | -89:42:30.30 | 12.69 | 54.133 (15) | 4668.37 (20) | 38 (2) | EB |
| 002-003056 | 18:36:30.712 | -89:11:02.84 | 12.76 | 75.57626 (20) | 4629.9895 (60) | 30 (2) | EA |
| 001-001090 | 20:37:01.418 | -89:44:34.88 | 12.82 | 23.7052 (70) | 4632.38 (10) | 10 (1) | EW |
| 001-000120 | 02:26:38.190 | -89:31:31.72 | 14.45 | 81.051070 (20) | 5058.9223 (24) | 335 (5) | EA |
| 001-000682 | 12:27:06.862 | -89:48:58.91 | 14.59 | 1.048938 (30) | 4627.2738 (97) | 28.0 (0.8) | EB |

Notes.

∗ This object was reported as a rotating ellipsoidal variable by Wang et al. (2015).
∗∗ This object was reported as a variable of unknown type by Wang et al. (2011) and was not identified by Wang et al. (2013).
† This object is EN Oct: we improved its period and eclipse depth compared to previous studies by Knigge & Bauernfeind (1967); Otero (2004).
†† This object is discussed in Sect. 7.2.5.