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Dependence comparisons of order statistics in the proportional hazards model

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Abstract
Let $X_1, \ldots, X_n$ be mutually independent exponential random variables with distinct hazard rates $\lambda_1, \ldots, \lambda_n$ and let $Y_1, \ldots, Y_n$ be a random sample from the exponential distribution with hazard rate $\bar{\lambda} = \sum_{i=1}^{n} \lambda_i/n$. Also let $X_{1:n} < \cdots < X_{n:n}$ and $Y_{1:n} < \cdots < Y_{n:n}$ be their associated order statistics. It is proved that for $1 \leq i < j \leq n$, the generalized spacing $X_{j:n} - X_{i:n}$ is more dispersed than $Y_{j:n} - Y_{i:n}$ according to dispersive ordering and for $2 \leq i \leq n$, the dependence of $X_{i:n}$ on $X_{1:n}$ is less than that of $Y_{i:n}$ on $Y_{1:n}$, in the sense of the more stochastically increasing ordering. This dependence result is also extended to the proportional hazard rates (PHR) model. This extends the earlier work of Genest et al. [(2009)]. On the range of heterogeneous samples. Journal of Multivariate Analysis 100: 1587–1592 who proved this result for $i = n$.

1. Introduction

Several notions of monotone dependence have been introduced and studied in the literature. Researchers have also developed the corresponding dependence (partial) orderings which compare the degree of (monotone) dependence within the components of different random vectors of the same length. For details, see the pioneering paper of Lehmann [16] and Chapter 5 of Barlow and Proschan [2] for different notions of positive dependence, and that of Kimeldorf and Sampson [12] for a unified treatment of families, orderings and measures of monotone dependence. For more detailed discussion of these concepts, see Chapter 2 of Joe [9] and Chapter 5 of Nelsen [17].

Let $\{X_1, \ldots, X_n\}$ be a set of continuous random variables. Many authors have investigated the nature of the dependence that may exist between two order statistics $X_{i:n}$ and $X_{j:n}$ for $1 \leq i < j \leq n$ under different distributional scenarios. It seems natural to expect some degree of positive dependence between them. It is well known that the order statistics based on independent (but not necessarily identically distributed) random variables are associated (cf., [2, p. 32]). This yields many useful product inequalities for order statistics of independent random variables, and in particular, $\text{Cov}(X_{i:n}, X_{j:n}) \geq 0$ for all $i$ and $j$, which was initially proved by Bickel [3] when the $X$’s are independent and identically distributed (i.i.d.).

Boland et al. [4] studied the problem of dependence among order statistics in detail and discussed different types of dependence that hold between them. It is shown in that paper that in the case of i.i.d. observations, any pair of order statistics is $TP_2$ dependent (also called monotone likelihood ratio dependent) which is the strongest type of dependence in the hierarchy of various dependence criteria as described in [2]). But this is not the case in the non-i.i.d. case. However, as is shown in Boland et al. [4], in general, whereas $X_{j:n}$ may not be stochastically increasing in $X_{i:n}$ for $i < j$, $X_{j:n}$ is always right tail increasing (RTI) in $X_{i:n}$. The reader is referred to Chapter 5 of Barlow and Proschan [2] for basic definitions and relations among various types of dependence.

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It is also of interest to compare the strength of dependence that may exist between two pairs of order statistics. When the parent distribution from which the random sample is drawn has an increasing hazard rate and a decreasing reverse hazard rate, Tukey [22] showed that

$$\text{Cov}(X_{i:n}, X_{j:n}) \leq \text{Cov}(X_{i:n}, X_{j:n})$$

(1.1)

for either $i = i'$ and $j \leq j'$; or $j = j'$ and $i' \leq i$. Kim and David [11] proved that if both the hazard and the reverse hazard rates of the $X_i$’s are increasing, then inequality (1.1) remains valid when $i = i'$ and $j \leq j'$; However, the inequality (1.1) is reversed when $j = j'$ and $i' \leq i$.

Let $X_1, \ldots, X_n$ be mutually independent exponentials with distinct hazard rates $\lambda_1, \ldots, \lambda_n$ and let $Y_1, \ldots, Y_n$ form a random sample from the exponential distribution with hazard rate $\bar{\lambda} = (\sum_{i=1}^n X_i)/n$. Sathe [18] proved that for any $i \in \{2, \ldots, n\}$

$$\text{corr}(X_{1:n}, X_{i:n}) \leq \text{corr}(Y_{1:n}, Y_{i:n}).$$

(1.2)

Although this observation is interesting, it merely compares the relative degree of linear association within the two pairs. It is now widely recognized, however, that margin-free measures of association are more appropriate than Pearson’s correlation, because they are based on the unique underlying copula which governs the dependence between the components of a continuous random pair.

In this paper, a long-standing open problem that for $1 \leq i < j \leq n$, the generalized spacing $X_{j:n} - X_{i:n}$ is more dispersed than $Y_{j:n} - Y_{i:n}$ according to dispersive ordering, has been solved (cf. [23]). This result is used to solve a related open problem that for $2 \leq i \leq n$, the dependence of $X_{i:n}$ on $X_{1:n}$ is less than that of $Y_{i:n}$ on $Y_{1:n}$, in the sense of the more stochastically increasing ordering. This dependence result is also extended to the proportional hazard rates (PHR) model. This extends the earlier work of Genest et al. [8] who proved this result for $i = n$. Section 2 is on some preliminaries where various definitions and notations are given. The main results of this paper are given in Section 3.

2. Preliminaries

For $i = 1, 2$, let $(X_i, Y_i)$ be a pair of continuous random variables with joint cumulative distribution function $H_i$ and margins $F_i, G_i$. Let

$$C_i(u, v) = H_i\{F_i^{-1}(u), G_i^{-1}(v)\}, \quad u, v \in (0, 1)$$

be the unique copula associated with $H_i$. In other words, $C_i$ is the distribution of the pair $(U_i, V_i) \equiv (F_i(X_i), G_i(Y_i))$ whose margins are uniform on the interval $(0, 1)$. See, for example, Chapter 1 of Nelsen [17] for details.

The most well-understood partial order to compare the strength of dependence within two random vectors is that of positive quadrant dependence (PQD) as defined below.

**Definition 2.1.** A copula $C_1$ is said to be less dependent than copula $C_2$ in the positive quadrant dependence (PQD) ordering, denoted $(X_1, Y_1) \prec_{\text{PQD}} (X_2, Y_2)$, if and only if

$$C_1(u, v) \leq C_2(u, v), \quad u, v \in (0, 1).$$

This condition implies that $\kappa(S_1, T_1) \leq \kappa(S_2, T_2)$ for all concordance measures meeting the axioms of Scarsini [19] like Kendall’s $\tau$ and Spearman’s $\rho$. See Tchen [21] for details.

Lehmann [16] in his seminal work introduced the notion of monotone regression dependence (MRD) which is also known as stochastic increasingness (SI) in the literature.
Definition 2.2. Let \((X, Y)\) be a bivariate random vector with joint distribution function \(H\). The variable \(Y\) is said to be stochastically increasing (SI) in \(X\) if for all \((x_1, x_2) \in \mathbb{R}^2\),

\[
x_1 < x_2 \Rightarrow P(Y \leq y \mid X = x_2) \leq P(Y \leq y \mid X = x_1), \quad \text{for all } y \in \mathbb{R}.
\]  

(2.1)

If we denote by \(H_x\) the distribution function of the conditional distribution of \(Y\) given \(X = x\), then (2.1) can be rewritten as

\[
x_1 < x_2 \Rightarrow H_{x_2} \circ H^{-1}_{x_1}(u) \leq u, \quad \text{for } 0 \leq u \leq 1.
\]  

(2.2)

Note that in case \(X\) and \(Y\) are independent, \(H_{x_2} \circ H^{-1}_{x_1}(u) = u\), for \(0 \leq u \leq 1\) and for all \((x_1, x_2)\). The SI property is not symmetric in \(X\) and \(Y\). It is a very strong notion of positive dependence and many of the other notions of positive dependence follow from it.

Denoting by \(\xi_p = F^{-1}(p)\), the \(p\)th quantile of the marginal distribution of \(X\), we see that (2.2) will hold if and only if for all \(0 \leq u \leq 1\),

\[
0 \leq p < q \leq 1 \Rightarrow H_{\xi_q} \circ H^{-1}_{\xi_p}(u) \leq u.
\]

In his book, Joe [9] mentions a number of bivariate stochastic ordering relations. One such notion is that of greater monotone regression (or more SI) dependence, originally considered by Yanagimoto and Okamoto [24] and later extended and further investigated by Schriever [20], Capéraà and Genest [5], and Avérous et al. [1], among others.

As discussed in the books by Joe [9] and Nelsen [17], a reasonable way to compare the relative degree of dependence between two random vectors is through their copulas. We will discuss here the degree of dependence between two random vectors in the sense of monotone regression dependence (stochastic increasingness).

Suppose we have two pairs of continuous random variables \((X_1, Y_1)\) and \((X_2, Y_2)\) with joint cumulative distribution functions \(H_i\) and marginals \(F_i\) and \(G_i\) for \(i = 1, 2\).

Definition 2.3. \(Y_2\) is said to be more stochastically increasing in \(X_2\) than \(Y_1\) in \(X_1\), denoted by \((Y_1 \mid X_1) \prec_{SI} (Y_2 \mid X_2)\) or \(H_1 \prec_{SI} H_2\), if

\[
0 < p \leq q < 1 \implies H_{2, \xi_{pq}} \circ H^{-1}_{2, \xi_{ip}}(u) \leq H_{1, \xi_{pq}} \circ H^{-1}_{1, \xi_{ip}}(u),
\]

(2.3)

for all \(u \in (0, 1)\), where for \(i = 1, 2\), \(H_{i,s}\) denotes the conditional distribution of \(Y_i\) given \(X_i = s\), and \(\xi_{ip} = F^{-1}_i(p)\) stands for the \(p\)th quantile of the marginal distribution of \(X_i\).

Obviously, (2.3) implies that \(Y_2\) is stochastically increasing in \(X_2\) if \(X_1\) and \(Y_1\) are independent. It also implies that if \(Y_1\) is stochastically increasing in \(X_1\), then so is \(Y_2\) in \(X_2\); and conversely, if \(Y_2\) is stochastically decreasing in \(X_2\), then so is \(Y_1\) in \(X_1\). The above definition of more SI is copula based and

\[H_1 \prec_{SI} H_2 \Rightarrow H_1 \prec_{PQD} H_2.\]

Avérous et al. [1] have shown in their paper that in the case of i.i.d. observations, the copula of any pair of order statistics is independent of the distribution of the parent observations as long as they are continuous. In a sense, their copula has a distribution-free property. But this is not the case if the observations are not i.i.d.

The natural question is to see if we can extend the result given in (1.2) to a copula-based positive dependence ordering. Genest et al. [8] proved that under the given conditions,

\[(X_{n:n} \mid X_{1:n}) \prec_{SI} (Y_{n:n} \mid Y_{1:n}).\]
It has been an open problem to see whether this result can be generalized from the the largest order statistic to other order statistics. We prove in this paper that for \(2 \leq i \leq n\),
\[
(X_{i:n} | X_{1:n}) \preceq_{SI} (Y_{i:n} | Y_{1:n}).
\]
This implies in particular that
\[
\kappa(X_{i:n}, X_{1:n}) \leq \kappa(Y_{i:n}, Y_{1:n}),
\]
where \(\kappa(S, T)\) represents any concordance measure between random variables \(S\) and \(T\) in the sense of Scarsini [19], for example, Spearman’s rho or Kendall’s tau. A related work to this problem is Dolati et al. [7].

3. Main results

The proof of our main result relies heavily on the notion of dispersive ordering between two random variables \(X\) and \(Y\), and properties thereof. For completeness, the definition of this concept is recalled below.

**Definition 3.1.** A random variable \(X\) with distribution function \(F\) is said to be less dispersed than another variable \(Y\) with distribution \(G\), written as \(X \leq_{\text{disp}} Y\) or \(F \leq_{\text{disp}} G\), if and only if
\[
F^{-1}(v) - F^{-1}(u) \leq G^{-1}(u) - G^{-1}(v)
\]
for all \(0 < u \leq v < 1\).

The dispersive order is closely related to the star-ordering which is a partial order to compare the relative aging or skewness and is defined as below.

**Definition 3.2.** A random variable \(X\) with distribution function \(F\) is said to be star ordered with respect to another random variable \(Y\) with distribution \(G\), written as \(X \leq_{\star} Y\) or \(F \leq_{\star} G\), if and only if
\[
\frac{G^{-1} \circ F(x)}{x}
\]
is increasing in \(x\).

For nonnegative random variables, the star order is related to the dispersive order by
\[
X \leq_{\star} Y \iff \log X \leq_{\text{disp}} \log Y.
\]

The proof of the next lemma, which is a refined version of a result by Deshpande and Kochar [6], on relation between the star order and the dispersive order, can be found in Kochar and Xu [15].

**Lemma 3.1.** Let \(X\) and \(Y\) be two random variables. If \(X \leq_{\star} Y\) and \(X \leq_{\text{st}} Y\), then \(X \leq_{\text{disp}} Y\).

The following result of Khaledi and Kochar [10] will be used to prove our main theorem.

**Lemma 3.2 ([10]).** Let \(X_i\) and \(Y_i\) be independent random variables with distribution functions \(F_i\) and \(G_i\), respectively for \(i = 1, 2\). Then
\[
X_2 \leq_{\text{disp}} X_1 \text{ and } Y_1 \leq_{\text{disp}} Y_2 \Rightarrow (X_2 + Y_2) | X_2 \preceq_{SI} (X_1 + Y_1) | X_1.
\]

The next theorem on dispersive ordering between general spacings which is also of independent interest, will be used to prove our main result.
Theorem 3.1. Let $X_1, \ldots, X_n$ be mutually independent exponential random variables with distinct hazard rates $\lambda_1, \ldots, \lambda_n$ and let $Y_1, \ldots, Y_n$ be a random sample from the exponential distribution with hazard rate $\bar{\lambda}$. Then for $1 \leq i < j \leq n$,

$$ (Y_{j:n} - Y_{i:n}) \leq_{\text{disp}} (X_{j:n} - X_{i:n}). $$

Proof. Yu [25] proved in Corollary 1 of his paper that for $1 \leq i < j \leq n$,

$$ (Y_{j:n} - Y_{i:n}) \leq_{\ast} (X_{j:n} - X_{i:n}). $$

and Kochar and Rojo [14] proved in their Corollary 2.1 that

$$ (Y_{j:n} - Y_{i:n}) \leq_{\text{distr}} (X_{j:n} - X_{i:n}). $$

Using (3.2) and (3.3), we get (3.1) using Lemma 3.1. □

Xu and Balakrishnan [23] proved a special case of the above result when $i = n$. Now, we give the main result of this paper.

Theorem 3.2. Let $X_1, \ldots, X_n$ be mutually independent exponentials with distinct hazard rates $\lambda_1, \ldots, \lambda_n$ and let $Y_1, \ldots, Y_n$ form a random sample from the exponential distribution with hazard rate $\bar{\lambda} = (\sum_{i=1}^{n} \lambda_i)/n$. Then, for $i \in \{2, \ldots, n\}$,

$$ (X_{i:n} | X_{1:n}) \leq_{\text{SI}} (Y_{i:n} | Y_{1:n}). $$

Proof. It follows from Theorem 3.1 above that for $2 \leq i \leq n$,

$$ (Y_{i:n} - Y_{1:n}) \leq_{\text{disp}} (X_{i:n} - X_{1:n}). $$

Kochar and Korwar [13] proved that $(X_{i:n} - X_{1:n})$ is independent of $X_{1:n}$ and $X_{1:n} \overset{\text{st}}{=} Y_{1:n}$. Similarly, $(Y_{i:n} - Y_{1:n})$ is independent of $Y_{1:n}$. Moreover,

$$ X_{1:n} \overset{\text{st}}{=} Y_{1:n}. $$

Now, we can express $X_{i:n}$ and $Y_{i:n}$ as

$$ X_{i:n} = (X_{i:n} - X_{1:n}) + X_{1:n} \quad \text{and} \quad Y_{i:n} = (Y_{i:n} - Y_{1:n}) + Y_{1:n}. $$

Using (3.4) and the above two equations, it follows from Lemma 3.2 that

$$ (X_{i:n} | X_{1:n}) \leq_{\text{SI}} (Y_{i:n} | Y_{1:n}), \quad \text{for } 2 \leq i \leq n. $$

This proves the required result. □

Remark 3.1. Since the copula of the order statistics of a random sample has the distribution-free property, it is not required that the common hazard rate of $Y$’s is necessarily $\bar{\lambda}$ in the above theorems. In fact, the $Y$’s could be a random sample from any continuous distribution.

The above theorem can be extended to the PHR model using the technique used in Genest et al. [8].

Theorem 3.3. Let $X_1, \ldots, X_n$ be independent continuous random variables following the PHR model with $\lambda_1, \ldots, \lambda_n$ as the proportionality parameters. Let $Y_1, \ldots, Y_n$ be i.i.d. continuous random variables with common survival function $F^1$, then

$$ (X_{i:n} | X_{1:n}) \leq_{\text{SI}} (Y_{i:n} | Y_{1:n}) \quad \text{for } 2 \leq i \leq n. $$
Proof. Let \( R(t) = -\log \{ \bar{F}(t) \} \) in the PHR model. Make the monotone transformations \( X_i^* = R(X_i) \) and \( Y_i^* = -\log \{ \bar{Y}(Y_i) \} \). Then, the transformed variable \( X_i^* \) has exponential distribution with hazard rate \( \lambda_i \), \( i = 1, \ldots, n \) and \( Y_i^* \) is a random sample from an exponential distribution with parameter \( \lambda \). Let \( X_1^* < \cdots < X_{(n)}^* \) and \( Y_1^* < \cdots < Y_{(n)}^* \) be the order statistics corresponding to the new sets of variables.

In view of their invariance by monotone increasing transformations of the margins, the copulas associated with the pairs \((X_{1:n},X_{2:n})\) and \((X_{1:n},X_{2:n})\) coincide. Similarly, the pairs \((Y_{1:n},Y_{2:n})\) and \((Y_{1:n},Y_{2:n})\) have the same copula.

Also, since the more SI dependence ordering is copula-based,

\[
(X_{n:n} \mid X_{1:n}) <_{SI} (Y_{n:n} \mid Y_{1:n}) \iff (X_{n:n}^* \mid X_{1:n}^*) <_{SI} (Y_{n:n}^* \mid Y_{1:n}^*).
\]

Under the conditions of Theorem 3.3, an upper bound on \( \kappa(X_{1:n},X_{i:n}) \) is given by \( \kappa(Y_{1:n},Y_{i:n}) \).

Avérous et al. [1] obtained an analytic expression for computing the exact values of the Kendall’s \( \tau \) for any pair of order statistics from a random sample from a continuous distribution which in our case reduces to

\[
\tau(Y_{1:n},Y_{i:n}) = 1 - \frac{2(n - 1)}{2n - 1} \left( \sum_{i=0}^{n} \binom{n - i}{s} \left( \frac{n - 2}{n - i + s} \right) \right).
\]

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