Preoperative beta-blocker in ventricular dysfunction patients: need a more granular quality metric

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Abstract

Background: The use of preoperative beta-blockers has been accepted as a quality standard for patients undergoing coronary artery bypass graft (CABG) surgery. However, conflicting results from recent studies have raised questions concerning the effectiveness of this quality metric. We sought to determine the influence of preoperative beta-blocker administration before CABG in patients with left ventricular dysfunction.

Methods: The authors analyzed all cases of isolated CABGs in patients with left ventricular ejection fraction less than 50%, performed between 2012 January and 2017 June, at 94 centres recorded in the China Heart Failure Surgery Registry database. In addition to the use of multivariate regression models, a 1–1 propensity scores matched analysis was performed.

Results: Of 6116 eligible patients, 61.7% received a preoperative beta-blocker. No difference in operative mortality was found between two cohorts (3.7% for the non-beta-blockers group vs. 3.0% for the beta-blocker group; adjusted odds ratio [OR] 0.82 [95% CI 0.58–1.15]). Few differences in the incidence of other postoperative clinical end points were observed as a function of preoperative beta-blockers except in stroke (0.7% for the non-beta-blocker group vs. 0.3 for the beta-blocker group; adjusted OR 0.39 [95% CI 0.16–0.96]). Results of propensity-matched analyses were broadly consistent.

Conclusions: In this study, the administration of beta-blockers before CABG was not associated with improved operative mortality and complications except the incidence of postoperative stroke in patients with left ventricular dysfunction. A more granular quality metric which would guide the use of beta-blockers should be developed.

Keywords: Beta-blocker, Quality metric, Coronary artery bypass grafting

Background

Coronary artery bypass grafting (CABG) is an essential therapeutic approach to reduce mortality and morbidity in patients with complex, multivessel coronary artery disease (CAD) [1, 2]. The perioperative medical treatment is key in the optimal success of the cardiac surgery. Clinical guidelines have recommended that beta-blockers should be used in heart failure and CABG patients without compelling contraindications [3, 4]. In the late 1990s, a large retrospective analysis demonstrated benefits of preoperative beta-blockers usage [5]. Different mechanisms may contributes to the protective effect of beta-blockers, for example, improvement in the oxygen supply–demand balance of the myocardium, decreasing effect of sympathetic nervous activity, suppression of...
dysrhythmias and remodelling of the left ventricular [6]. Our group previously reported that postoperative beta-blocker therapy was associated a lower risk of long-term mortality and adverse cardiovascular events [7]. Since 2007, the use of preoperative beta-blockers has been used as a quality standard for patients undergoing CABG [8].

However, controversies still exist in the literature regarding the effectiveness of beta-blockers for patients undergoing CABG [9–12]. These studies failed to show perioperative mortality advantage in patients receiving beta-blockers before surgery. A recently meta-analysis suggests that the use of preoperative beta-blockers did not reduce either operative mortality or the incidence of postoperative complications [13]. Therefore, the preoperative use of beta-blockers as a quality indicator has been questioned.

To date, the efficacy and safety of beta-blocker use during the preoperative period of CABG have not been adequately evaluated in patients with left ventricular dysfunction. Accurately identifying and utilizing measures of performance and quality is essential. Prospective randomized clinical trials to investigate the effect of preoperative beta-blocker use on cardiovascular patients is difficult and there are concerns that the design of such trials would entail an unacceptable risk for the withdrawal of beta-blockers in patients already taking this medication can lead to substantial morbidity and even mortality [10].

As large-scale registries may overcome these difficulties to support clinical decisions, the current study aims to review a large national database to assessed whether preoperative administration of beta-blockers was associated with improved early clinical outcomes after CABG in patients with left ventricular dysfunction.

Methods
Study design
The China Heart Failure Surgery Registry (China-HFSR) was led by Fuwai Hospital and other representative cardiac centres in different regions around China. In total, 94 centres with annual surgery volumes > 100 were included as participants in the study. We included patients ≥17 years old who underwent CABG from January 2012 to June 2017 with documented LVEF < 50%. Patients were excluded if they underwent concomitant valve or other surgeries or non-elective surgeries. We also excluded patients who have preoperative intra-aortic balloon pump insertion, cardiogenic shock and third degree heart block (Fig. 1). These patients were then stratified according to preoperative beta-blocker administration. All CABG procedures represented standard surgical approaches to surgical myocardial revascularization with and without the use of cardiopulmonary bypass support. This study was approved by the institutional review board at Fuwai Hospital (approval number 887, April 25th, 2017) and carried out in accordance with relevant guidelines and regulations. The informed consent was provided by participants.

Data collection
All data were collected at the local sites from the medical records. The requirements for data collection and the definitions of variables were clearly identified. Standardized electronic case report forms were completed at the local sites and then submitted online to the data processing centre. All data were into the database separately by two trained technicians. Two separate reviewers from the data processing centre randomly selected and assessed 5–10% of each of the participating centres’ medical records during annual on-site audits. We compared the data in the database and the original medical records. A committee composed of physicians and surgeons determined the correct final value when there was a disagreement. In all patients included in China-HFSR database, 90 (1.5%) patients were without listed height, 74 (1.2%) without listed weight and 1 without data regarding smoking history. Considering the fact that they only accounted for a very small proportion of our patients, we imputed missing continuous variables (height and weight) with different mean values for the sexes. The missing categorical variable (smoking history) was imputed with negative values. So that patients who might have experienced the end point (in-hospital death) would not be excluded from analysis simply for 1 or 2 missing variables among the many examined.

Clinical data
The preoperative variables included age, gender, body mass index, smoking history, New York Heart Association (NYNH) classification, Canadian Cardiovascular Society (CCS) classification, diabetes mellitus (DM), hypertension, hyperlipidemia, renal failure, chronic obstructive pulmonary disease (COPD), cerebrovascular accident, carotid disease and other peripheral arterial disease, preoperative atrial fibrillation, previous myocardial infarction (MI), percutaneous transluminal coronary angioplasty (PTCA) history, No. of diseased vessels, Left main CAD, LVEF, preoperative creatinine and prior cardiovascular surgeries.

The major postoperative complications included reintubation, MI, mediastinal infection, stroke, renal failure, multiple organ dysfunction syndrome, postoperative atrial fibrillation and reoperation for bleeding. MI was counted as a complication if it newly occurred postoperatively and was defined as any one of the following: MI documented in the medical record with an elevation of
cardiac troponin values with at least one value above the 10 times 99th percentile upper reference limit or electrocardiograph documented ST-segment elevation in evolution, Q waves 0.03 s in width and/or one-third or greater of the total QRS complex in 2 or more contiguous leads, or new left bundle branch block [14]. Mediastinal infection was defined according to the expert consensus [15]. Stroke was defined as a central neurological deficit persisting >24 h (i.e., extremity weakness or loss of motion, loss of consciousness, loss of speech, visual field cuts). Renal failure was defined as an increase in serum creatinine level to >4 mg/dL, 3× the most recent preoperative creatinine level, or a new postoperative need for dialysis. Reoperation for bleeding was defined as chest tube drainage ≥200 mL/h for at least 3 h requiring surgical intervention.

**Statistical analysis**
Continuous variables are expressed as either mean±standard deviation or medians and quartiles depending upon overall variable distribution. Categorical variables are presented as frequencies and percentages. We performed a t-test for normally distributed continuous variables; otherwise, the Mann–Whitney U test or Kruskal–Wallis H test was used. Chi square tests or Fisher’s exact tests were used for categorical variables. Cochran-Armitage trend test were used to examine the trend of beta-blocker use during the study period.

We used the following 2 techniques to adjust for selection bias when comparing outcomes of the beta-blocker versus non-beta-blocker groups: multiple logistic regression modelling and propensity matching. For the regression-based analyses, the association between preoperative beta-blocker use and each clinical end point were adjusted for clinical end point were adjusted for baseline patient risk by inclusion of the following validated and widely accepted measures of patient-level covariates: age, body mass index, sex, smoking history, diabetes mellitus, hypertension, hyperlipidemia, chronic obstructive pulmonary disease, cerebrovascular accident, previous MI, PTCA history, LVEF, preoperative creatinine, CCS classification, NYHA classification, No. of diseased vessels, Left main CAD, preoperative atrial fibrillation and prior cardiovascular surgery history. Each logistic model also included year of surgery and a set of fixed-effect hospital-specific intercept variables [16]. Model results are reported as odds ratios (OR) with a 95% confidence interval.

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**Fig. 1** Patient flowchart. CABG, coronary artery bypass grafting; IABP, intra-aortic balloon pump

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China-HFSR database, CABG cases, January 2012 to June 2017
N=8,316

Concomitant valve or other surgeries excluded
N=1,785

Isolated CABG cases
N=6,531

415 excluded
Cardiogenic shock 32
Third degree heart block 17
Preoperative IABP Insertion 182
Urgent or emergency surgery 184

Cases enrolled
N=6,116
The second method of adjusting for selection bias involved matching patients with similar estimated probability of receiving beta-blockers (propensity score). The propensity score was calculated by a multivariable logistic regression model which was developed using the same covariates listed above for the regression-based analyses. Then we matched patients in a 1:1 fashion without replacement [17]. ORs with 95% CIs comparing the frequency of each end point for patients receiving vs not receiving beta-blockers were estimated using univariable logistic regression.

Additional analysis were performed to examine whether the association between beta-blockers and mortality differed across prespecified subgroups based on age, sex, ejection fraction, diabetes mellitus, hypertension and chronic lung disease. Subgroup-specific ORs were estimated and displayed with 95% CIs.

All reported \( P \) values are 2 sided, and values of \( P < 0.05 \) were considered to indicate statistical significance. All statistical analysis was performed using SPSS version 22.0 (IBM Corp., Armonk, NY, USA).

Results

Patients characteristics

Of 6,116 patients who met study inclusion criteria, 61.7% of patients received a preoperative beta-blocker. During the study period, no significant trend was found in the beta-blocker use (Additional file 1: Figure S1, \( P = 0.163 \)). According to the preoperative profiles, women accounted for 16.2% of the patients, and 34.0% of the overall patient population had diabetes mellitus; the mean patient age was 61.3 ± 9.2 years. Small differences in baseline characteristics existed between study groups. Patients receiving beta-blockers were more likely to have DM, hypertension prior PTCA history and left main CAD while hyperlipidemia, COPD, Carotid disease and triple vessel disease were less common in these patients (Table 1).

Operative outcomes

Table 2 summarizes the outcomes from the unmatched groups. No difference in mortality was found between two cohorts (3.7% for the non-beta-blockers group vs 3.0% for the beta-blocker group; adjusted OR 0.82 [95% CI 0.58–1.15]). Few differences in the incidence of other postoperative clinical end points were observed as a function of preoperative beta-blockers except in the incidence of stroke (0.7% for the non-beta-blocker group vs. 0.3 for the beta-blocker group; adjusted OR 0.39 [95% CI 0.16–0.96]).

After propensity matching, comparable groups of 2430 each were created (Additional file 3: Table S1; Additional file 2: Figure S2). Outcomes for propensity-matched patients receiving vs not receiving beta-blockers are displayed in Table 3. Operative mortality was similar for the 2 groups (3.7% vs. 3.0% for the beta-blocker vs non-beta-blocker groups; OR 0.85 [95% CI 0.62–1.16]). Less frequency of stroke was found in patients receiving preoperative beta-blockers (0.3%) compared with non-beta-blockers (0.7%; OR 0.41 [95% CI 0.17–0.99]). Of other clinical end points, no significant differences were found.

Prespecified subgroup analysis

Figure 2 illustrates the effects of beta-blocker therapy among prespecified patient subgroups in the propensity matched cohort. For each subgroups, all the calculations include 1.0 in the 95% CI for the OR and the interaction \( P \) value was not significant (\( P \geq 0.05 \)).

Discussion

The present study reports upon the effect of preoperative beta-blocker use on the operative outcomes of CABG. We performed analyses on the unmatched and propensity-matched cohorts, controlling for the preoperative risk factors. No statistical mortality benefit was associated with the use of preoperative beta-blockers. This finding was consistent among the various subgroups. Other end points were equal between groups except for postoperative stroke, which incidence was higher in the non-beta-blocker group. These results re-examined the appropriateness of preoperative beta-blocker use as a useful quality measure for isolated CABG in patients with left ventricular dysfunction.

As beta-blockers have been proved to be effective in correcting the imbalance between oxygen demand and supply, this medication has been used routinely as a main therapy for patients with cardiovascular disease over the past 2 decades. The basis for preoperative beta-blocker use was extrapolated from studies in noncardiac surgical patient populations and a single nationwide retrospective analysis from the late 1990s [5, 18]. At present, clinical guidelines for CABG surgery recommend preoperative beta-blockers for patients without contraindications [1, 4, 19]. However, there always has been controversy regarding whether preoperative beta-blocker use should be used as an quality metric, on account of it simply identifies whether a patient either received or did not received a preoperative beta-blocker.

The large observational analysis reported by Ferguson and colleagues revealed a slightly lower mortality for patients undergoing CABG [5]. However, this benefit was limited in this analysis as patients with LVEF less than 30% were associated with a trend toward an increased mortality rate. In 2003, a retrospective review of propensity-matched CABG patients reported by Srinivasan et al. revealed that preoperative beta-blocker therapy was not associated with differences in operative outcomes.
[20]. Brinkman and colleagues similarly failed to show in their series propensity-matched studies that mortality or major complications benefits for preoperative beta-blocker use [11, 12]. In 2013, Lapar et al. reported potential negative impact of preoperative beta-blocker therapy [9]. They found association between perioperative myocardial infarction and preoperative beta-blocker therapy and observed patients treated with preoperative beta-blockers more commonly underwent intraoperative blood product transfusions. A recent meta-analysis including 6 observational studies with a total of 1,231,850 patients showed that preoperative beta-blocker use did not significantly reduce operative mortality and postoperative complications but significantly increased the incidence of atrial fibrillation [13].

In the present study, preoperative beta-blocker use in isolated CABG patients with left ventricular dysfunction was not associated with lower operative mortality in the

| Variables                           | All patient (n = 6116) | Beta-blocker use groups | P     |
|-------------------------------------|------------------------|-------------------------|-------|
|                                     | No (n = 2343)          | Yes (n = 3773)          |       |
| Age, mean (SD), years               | 61.3 (9.2)             | 61.4 (9.3)              | 61.3 (9.1) | 0.921 |
| Female, n (%)                       | 988 (16.2)             | 402 (17.2)              | 586 (15.5) | 0.093 |
| BMI, mean (SD)                      | 24.7 (3.2)             | 24.7 (3.1)              | 24.8 (3.2) | 0.445 |
| Smoking history, n (%)              | 3388 (55.4)            | 1304 (55.7)             | 2084 (55.2) | 0.748 |
| Diabetes mellitus, n (%)            | 2079 (34.0)            | 757 (32.3)              | 1322 (35.0) | 0.028 |
| Hypertension, n (%)                 | 3353 (54.8)            | 1237 (52.8)             | 2116 (56.1) | 0.012 |
| Hyperlipemia, n (%)                 | 1970 (32.2)            | 816 (34.8)              | 1154 (30.6) | 0.001 |
| Chronic renal failure, n (%)        | 113 (1.8)              | 37 (1.6)                | 76 (2.0) | 0.219 |
| COPD, n (%)                         | 94 (1.5)               | 46 (2.0)                | 48 (1.3) | 0.033 |
| Peripheral artery disease, n (%)    | 260 (4.3)              | 103 (4.4)               | 157 (4.2) | 0.658 |
| Carotid disease, n (%)              | 1000 (16.4)            | 441 (18.8)              | 559 (14.8) | 0.001 |
| Cerebrovascular accident, n (%)     | 527 (8.6)              | 191 (8.2)               | 336 (8.9) | 0.307 |
| Creatinine, median (25th, 75th percentile), umol/dL | 82.0 (70.0,96.6) | 81.7 (69.0,95.9) | 44.0 (40.46) | 0.132 |
| Left main CAD, n (%)                | 1625 (26.6)            | 589 (25.1)              | 1036 (27.5) | 0.046 |
| Triple vessel disease, n (%)        | 4218 (69.0)            | 2275 (73.5)             | 1443 (61.6) | <0.001 |
| Previous MI, n (%)                  | 2561 (41.9)            | 952 (40.6)              | 1609 (42.6) | 0.121 |
| PTCA history, n (%)                 | 760 (12.4)             | 255 (10.9)              | 505 (13.4) | 0.004 |
| CCS class                           |                        |                         |       |
| I, n (%)                            | 1181 (19.3)            | 508 (21.7)              | 673 (17.8) | 0.001 |
| II, n (%)                           | 886 (14.5)             | 348 (14.9)              | 216 (5.7) | 0.001 |
| III, n (%)                          | 2086 (34.1)            | 666 (28.4)              | 926 (24.5) | 0.001 |
| IV, n (%)                           | 1606 (26.3)            | 680 (29.0)              | 1420 (37.5) | 0.001 |
| LVF, Mean (SD), %                   | 42.3 (5.2)             | 42.3 (5.1)              | 42.3 (5.3) | 0.484 |
| LVEDD, Median (25th, 75th percentile), mm | 49 (55, 60)          | 55 (47, 60)             | 56 (50,60) | 0.002 |
| NYHA class                          |                        |                         |       |
| I, n (%)                            | 865 (14.1)             | 257 (11.0)              | 608 (16.1) | 0.001 |
| II, n (%)                           | 2252 (36.8)            | 864 (36.9)              | 1388 (36.8) | 0.001 |
| III, n (%)                          | 2694 (44.0)            | 1091 (46.6)             | 1603 (42.5) | 0.001 |
| IV, n (%)                           | 305 (5.0)              | 131 (5.6)               | 174 (4.6) | 0.001 |
| Atrial fibrillation, n (%)          | 123 (2.0)              | 44 (1.9)                | 79 (2.1) | 0.559 |
| Prior cardiovascular surgery, n (%) | 69 (1.1)               | 23 (1.0)                | 46 (1.2) | 0.393 |
| Preoperative ACEI/ARB, n (%)        | 1951 (29.9)            | 535 (21.3)              | 1416 (35.3) | <0.001 |
| STS PROM, Median (25th, 75th percentile),% | 3.0 [2.1, 4.4]  | 2.9 [2.0, 4.3]          | 3.1 [2.2, 4.6] | <0.001 |

ACEI, angiotensin-converting enzyme inhibitors; ARB, angiotensin receptor blockers; BMI, Body Mass Index; CAD, coronary vascular disease; CCS, Canadian Cardiovascular Society; COPD, chronic obstructive pulmonary disease; LVEDD, left ventricular end-diastolic dimension; LVF, left ventricular ejection fraction; MI, myocardial infarction; NA, not available; NYHA, New York Heart Association; PROM, predicted risk of mortality; PTCA, percutaneous transluminal coronary angioplasty; SD, standard deviation; STS, Society of Thoracic Surgeons
The whole study cohort (3.0% in beta-blockers group vs. 3.7% in non-beta-blockers groups; adjusted OR 0.82 [95% CI 0.58–1.15]) or in the propensity-matched cohort (3.2% in beta-blockers group vs. 3.7% in non-beta-blockers groups; OR 0.85 [95% CI 0.62–1.16]). We analyzed the association of beta-blocker use with operative outcomes in various subgroups but the results were similar in all cases.

In regard to postoperative complications, we were unable to demonstrate any benefit of preoperative beta-blocker use except postoperative stroke. The 30-day results of the PeriOperative ischemic Evaluation (POISE) trial showed a significant reduction in cardiac events at the cost of a significant increase in the incidence of total mortality and stroke [21]. Concerns were raised as postoperative stroke might be associated with perioperative hypotension related with beta-blocker therapy. However, to our knowledge, there is lack of evidence linking beta-blocker use to hypotension after CABG. The chronic use of beta-blockers may improve baroreflex sensitivity and lower heart rate variability [22]. Therefore, beta-blockers may prevent severe hemodynamic fluctuations after CABG thus reducing the incidence of stroke. Moreover, long-term beta-blocker use has been shown to slow artery plaque development and improve plaque stabilization [23, 24]. This in turn could result in less microembolizations during and after CABG and therefore less postoperative stroke.

Our report has important clinical and health policy related meanings. At present, the STS has identified preoperative beta-blocker use, which simply identifies whether a patient either received or did not received a preoperative beta-blocker, as a quality metric. Further details related to beta-blocker type, dose, timing of administration, goal heart rate, or duration of preoperative therapy is not involved in the assessment. Without more granular metric, the true efficacy of preoperative beta-blocker therapy will remain uncertain. As the existing binary metric is unlikely to show any difference, then it should not be a quality metric. A quality metric should

| Table 2 | Number of end point events and covariate-adjusted ORs in overall cohort |
|-----------------|-----------------|------------------|
| End point       | No. (%) of events by group | OR (95% CI)      |
|                 | No beta-blocker (n = 2343) | Beta-blocker (n = 3773) |
| Mortality       | 87 (3.7)          | 113 (3.0)        | 0.82 (0.58–1.15) | 0.256 |
| Re-intubation   | 60 (2.6)          | 91 (2.4)         | 1.03 (0.69–1.54) | 0.869 |
| Postoperative MI| 24 (1.0)          | 24 (0.6)         | 0.95 (0.44–2.07) | 0.894 |
| Mediastinal infection | 19 (0.8) | 28 (0.7)        | 0.92 (0.45–1.89) | 0.825 |
| Postoperative AF| 634 (27.1)        | 1089 (28.9)      | 1.07 (0.94–1.21) | 0.135 |
| Postoperative stroke | 17 (0.7) | 12 (0.3)        | 0.37 (0.16–0.96) | 0.040 |
| Postoperative renal failure | 43 (1.8) | 62 (1.6) | 0.85 (0.53–1.36) | 0.494 |
| MODS            | 43 (1.8)          | 62 (1.6)         | 0.99 (0.56–1.75) | 0.975 |
| Re-operation    | 37 (1.6)          | 90 (2.4)         | 1.31 (0.85–2.02) | 0.228 |

| Table 3 | Number of end point events and ORs in propensity-matched cohort |
|-----------------|-----------------|------------------|
| End point       | No. (%) of events by group | OR (95% CI)      |
|                 | No beta-blocker (n = 2340) | Beta-blocker (n = 2340) |
| Mortality       | 87 (3.7)          | 74 (3.2)         | 0.85 (0.62–1.16) | 0.298 |
| Re-intubation   | 60 (2.6)          | 58 (2.5)         | 0.97 (0.67–1.39) | 0.852 |
| Postoperative MI| 24 (1.0)          | 16 (0.7)         | 0.66 (0.35–1.25) | 0.207 |
| Mediastinal infection | 19 (0.8) | 20 (0.9)       | 1.05 (0.56–1.98) | 0.872 |
| Postoperative AF| 634 (27.1)        | 628 (26.8)       | 0.987 (0.87–1.12) | 0.843 |
| Postoperative stroke | 17 (0.7) | 7 (0.3)       | 0.41 (0.17–0.99) | 0.048 |
| Postoperative renal failure | 43 (1.8) | 41 (1.8) | 1.17 (0.75–1.81) | 0.499 |
| MODS            | 31 (1.3)          | 32 (1.4)         | 1.03 (0.63–1.70) | 0.899 |
| Re-operation    | 37 (1.6)          | 53 (2.3)         | 1.44 (0.94–2.20) | 0.090 |

AF, atrial fibrillation; CI, confidence interval; MI, myocardial infarction; MODS, multiple organ dysfunction syndrome; OR, odds ratio.
be one that makes a difference in outcomes. In order to really help guide health policy makers and guide surgeons in the future, we believe we need to look at preoperative beta-blocker use in greater depth. More granular in definitions of whether a goal preoperative beta-blocker effect was achieved is needed.

We suggest that the appropriate use of beta-blockers is where we need to take effort in the future investigation. For patients who are on beta-blockers, we advocate that these patients continue their preoperative beta-blockers therapy. Giving a patient a preoperative beta-blocker right before they go into the operation room to meet a predefined measure of cardiac surgical quality should be questioned. We assumed that a balance between the beta-blocker does to achieve a target heart rate in relation to the drug side effects should be optimized. There is a demand for further studies.

**Limitations**

Our study has several limitations. First, salient details related to the type of beta-blockers, dosage, timing and duration of beta-blocker therapy, heart rate, blood pressure and postoperative vasopressor cannot be determined from this data set. Different types of beta-blockers may not have the same effect. Clemente-Moragón and colleagues found metoprolol exerts a disruptive action on neutrophil dynamics during exacerbated inflammation, resulting in an infarct-limiting effect not observed with atenolol or propranolol in mouse models [25]. We believe it is of important to compare the different clinically approved beta-blockers in CABG patients with more granular data to improve the existing quality metric. Second, selection bias regarding the use of beta-blockers is unavoidable in observational studies. The propensity score used to adjust for baseline beta-blocker use can only account for measured covariates; thus we could not exclude the influence of unmeasured confounders on clinical outcomes. Finally, all analyses were limited to short-term outcomes. Any long-term benefit related to preoperative beta-blockers therapy would not be demonstrated.

**Conclusion**

In this study, the administration of beta-blockers before CABG in patients with left ventricular dysfunction was not associated with improved operative mortality and complications except the incidence of postoperative stroke. Simply identifying whether or not beta-blockers are used before surgery should not be used as a measure of surgical quality. Beta-blockers are an important and effective tool in the care of specific patients undergoing cardiac surgery in specific clinical scenarios. However, a more granular quality metric should be developed.

**Abbreviations**

CABG: Coronary artery bypass grafting; CAD: Coronary artery disease; CCS: Canadian Cardiovascular Society; CI: Confidence interval; COPD: Chronic obstructive pulmonary disease; DM: Diabetes mellitus; LVEF: Left ventricular ejection fraction; MI: Myocardial infarction; NYHA: New York Heart Association; OR: Odds ratio; PTCA: Percutaneous transluminal coronary angioplasty.

**Supplementary Information**

The online version contains supplementary material available at https://doi.org/10.1186/s12872-021-02371-1.

Additional file 1. Supplemental materials.

Additional file 2. Trend in beta-blocker use by year of operation.

Additional file 3. A, Preoperative characteristics varied widely between no beta-blockers group and beta-blockers group. B, After matching, there were no significant differences between the matched cohorts (no beta-blockers group vs beta-blockers group).

**Acknowledgements**

The authors thank all China-HFSR investigators and participants for their contributions: Xin Chen, Yingbin Xiao, Fei Li, Feng Wan, Tao Han, Liangwan Chen, Hongyu Liu, Su Liu, Zhaoyun Cheng, Guimin Zhang, Zhongmin Liu, Chunsheng Wang, Hushan Wang, Zhenya Shen, Zhigang Liu, Chengchao Sun, Xinmin Zhou, Yiming Ni.
Authors contributions
SH conceived of the China-HFSR study and take responsibility for all aspects of it. SH, HT and XH designed the study. HT wrote the first draft of the article and did the statistical analysis, with further contributions from KC, JH and SL. All authors read and approved the final manuscript.

Funding
The China-HFSR was supported by the 13th Five-Year National Science and Technology Major Project of China (2016YFC1300900). They had no role in designing the study, data collection or analysis, interpretation of data, or writing the manuscript.

Availability of data and materials
The datasets generated during and analyzed during the current study are not publicly available due to the datasets also forms part of other ongoing studies but are available from the corresponding author on reasonable request.

Declarations

Ethics approval and consent to participate
This study was approved by the institutional review board at Fuwai Hospital (approval number 887, April 25th, 2017) and carried out in accordance with relevant guidelines and regulations. The informed consent was signed by participants.

Consent for publication
Not applicable.

Competing interests
The authors declare that they have no competing interests.

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Received: 14 January 2021 Accepted: 28 October 2021

Published online: 19 November 2021

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