Population Pharmacokinetics of Telapristone (CDB-4124) and its Active Monodemethylated Metabolite CDB-4453, with a Mixture Model for Total Clearance

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Abstract. Telapristone is a selective progesterone antagonist that is being developed for the long-term treatment of symptoms associated with endometriosis and uterine fibroids. The population pharmacokinetics of telapristone (CDB-4124) and CDB-4453 was investigated using nonlinear mixed-effects modeling. Data from two clinical studies (n=32) were included in the analysis. A two-compartment (parent) one compartment (metabolite) mixture model (with two populations for apparent clearance) with first-order absorption and elimination adequately described the pharmacokinetics of telapristone and CDB-4453. Telapristone was rapidly absorbed with an absorption rate constant (Ka) of 1.26 h⁻¹. Moderate renal impairment resulted in a 74% decrease in Ka. The population estimates for oral clearance (CL/F) for the two populations were 11.6 and 3.34 L/h, respectively, with 25% of the subjects being allocated to the high-clearance group. Apparent volume of distribution for the central compartment (V2/F) was 37.4 L, apparent inter-compartmental clearance (Q/F) was 21.9 L/h, and apparent peripheral volume of distribution for the parent (V4/F) was 120 L. The ratio of the fraction of telapristone converted to CDB-4453 to the distribution volume of CDB-4453 (Fmet/est) was 0.20/L. Apparent volume of distribution of the metabolite compartment (V3/F) was fixed to 1 L and apparent clearance of the metabolite (CLM/F) was 2.43 L/h. A two-compartment parent-metabolite model adequately described the pharmacokinetics of telapristone and CDB-4453. The clearance of telapristone was separated into two populations and could be the result of metabolism via polymorphic CYP3A5.

KEY WORDS: CDB-4453; mixture model; parent-metabolite; population pharmacokinetics; telapristone (CDB-4124).

INTRODUCTION

Over the last 25 years, several progesterone receptor antagonists have been evaluated for the treatment of progesterone-dependent diseases such as endometriosis and uterine leiomyoma (1). However, many of the potential candidates also exhibited potent antiglucocorticoid activity (1,2). Additional anti-progestins have been synthesized and tested in order to identify candidates that are progesterone antagonist with little to no antiglucocorticoid activity. Telapristone (US adopted name: telapristone acetate) was one of the compounds developed by the Contraception and Reproductive Health Branch (National Institute of Child Health and Human Development) for the chronic treatment of symptoms associated with endometriosis and leiomyoma (3).

Telapristone opposes progesterone action at the molecular and receptor level, and has been shown in animals to exhibit potent anti-progesterone effects. Administration of telapristone to endometriosis-induced cynomolgus monkeys demonstrated that telapristone lead to reduction in the height and development of the endometrium as well as changes in the mitotic index, glands and stroma. It was also noted that telapristone was not associated with any antiglucocorticoid activity, as shown by a lack of measurable increase in serum cortisol. Clinical evaluation of hormone levels and ovulation showed that telapristone did not affect serum estrogen or progesterone and did not appear to affect ovulation when used once every 7 days.

A phase I/II study (ZN001) in pre-menopausal women (n=30) with symptomatic leiomyoma was completed (Trial identification number: CTNCT01069094). It consisted of a single-dose pharmacokinetic arm (telapristone only) followed in order to identify candidates that are progesterone antagonist with little to no antiglucocorticoid activity. Telapristone (US adopted name: telapristone acetate) was one of the compounds developed by the Contraception and Reproductive Health Branch (National Institute of Child Health and Human Development) for the chronic treatment of symptoms associated with endometriosis and leiomyoma (3).

Telapristone opposes progesterone action at the molecular and receptor level, and has been shown in animals to exhibit potent anti-progesterone effects. Administration of telapristone
bleeding days as well as a greater reduction in global tumor size (4).

The reduced antiglucocorticoid activity (5), indications that physiologic concentrations of estrogen are preserved and results from efficacy trials for both uterine fibroids and endometriosis (4,7–9), suggest that telapristone may offer advantages in both safety and efficacy that other currently available treatments do not.

Data in mice indicates that telapristone is primarily metabolized in the liver. It has been proposed that telapristone (Fig. 1) undergoes demethylation to produce mono (CDB-4453) and didemethylated metabolites as well as hydroxylation of the propynyl group to produce the hydroxylated metabolite. Like mifepristone, it is proposed that metabolism of telapristone to the monodemethylated and hydroxylated metabolite is rapid, but formation of the didemethylated metabolite occurs more slowly and to a much lesser extent (6,10,11). Out of the three characterized metabolites, in vitro data indicated that CDB-4453 was equipotent to CDB-4124 as an anti-progesterone. In addition, a study evaluating the in vitro activity of CDB-4453 showed that it had less antiglucocorticoid activity than telapristone (5,6). Since CDB-4453 and telapristone share similar anti-progestational activity in vitro, the presence of CDB-4453 might contribute significantly to the in vivo anti-progesterone activity of telapristone. Therefore, it is important to understand the kinetics of both telapristone and CDB-4453.

Information pertaining to the pharmacokinetics and sources of variability within parameter estimates for telapristone and CDB-4453 is limited. The aim of this analysis was to develop a pharmacokinetic model for telapristone and CDB-4453, evaluate the effects of covariates on the inter-individual variability associated with parameters, and simulate concentration time profiles of telapristone and CDB-4453 in order to assess the range of variability in pharmacokinetic parameters.

MATERIALS AND METHODS

Description of Studies

Study ZP-005 was a phase I/II, open-label study to evaluate the safety and pharmacokinetics of telapristone in female patients with moderate hepatic impairment (meeting the Child–Pugh class B severity criteria) versus healthy volunteers. Patients and volunteers received a single oral dose of 25 mg of telapristone acetate (fasted) followed by 50 mg of telapristone acetate (fasted) after a 14-day washout period (Trial identification number: NCT00741273).

Study ZP-006 was a phase I/II single-dose open-label parallel study to evaluate the pharmacokinetics and safety profile of telapristone in female patients with mild and moderate renal impairment versus healthy volunteers. Subjects received a single oral dose of 50 mg of telapristone acetate (fasted) (Trial identification number: NCT00787618).

Pharmacokinetic samples for both studies were taken at pre-dose, 0.25, 0.5, 0.75, 1, 1.5, 2, 2.5, 3, 5, 7, 9, 12, 16, 20, 24, 28, 32, 36, 40, 44, and 48 h postdose.

All protocols, amendments, and informed consent forms were reviewed and approved by an Institutional Review Board (IRB) at each study site. The study did not start until the IRB had approved the protocol or a modification. The IRB was constituted and operated in accordance with the principles and requirements described in the US Code of Federal Regulations (21 CFR Part 56).

![Fig. 1. Proposed metabolism of telapristone](image-url)
Bioanalytical Methods

Plasma samples were analyzed for both telapristone and CDB-4453 using LC/MS/MS. Sample analysis involved the extraction of telapristone or CDB-4453, and the internal standard mifepristone, using protein precipitation. This extract was then subjected to reverse-phase high-performance liquid chromatography on an Aquasil C18 column and detection of the analytes by tandem mass spectroscopy using the Sciex API3000 LC/MS/MS. The lower limit of quantification was 5 ng/mL. The coefficient of variation for in-assay precision and accuracy ranged from 1.5% to 4.7% and −1.0% to 5.5%, respectively. This method was previously validated over the range of 5.00–1,000 ng/mL with 100-μl aliquot volume.

Data Analysis

Population Pharmacokinetic Analysis

Data Handling. Due to the difference in molecular weight of telapristone (505.6 g/mol; http://www.ama-assn.org/resources/doc/usan/telapristone-acetate.pdf) and CDB-4453, concentrations of telapristone and CDB-4453 were converted to equivalent nanomoles per liter. Concentrations of telapristone and CDB-4453 were also log-transformed before the analysis. All concentration values below the quantification limit (BQL) of the assay were excluded from the pharmacokinetic analysis.

Base Pharmacokinetic Model. Non-linear mixed-effect model building was conducted using NONMEM® 7 (ICON Development Solutions, Ellicott City, MD, USA) (12) with G95 Fortran compiler (Free Software Foundation, Boston, MA, USA). All models were fitted using the first-order conditional estimation method. Output was processed using PDx-Pop 4.0 (ICON Development Solutions, Ellicott City, MD, USA) and Xpose version 4.0 (Uppsala University, Uppsala, Sweden) (13). Graphical plots were produced using S-PLUS version 8.1 (TIBCO, Somerville, MA, USA) and R 2.11.1 (Free Software Foundation, Boston, MA, USA). Model selection was guided by the plausibility of the estimates, minimum objective function value (MOFV), Akaike Information Criterion, condition number, visual inspection of diagnostic plots and the precision of parameter estimates.

Based on visual inspection of the concentration-time profiles of telapristone and CDB-4453, and previously published data (4), two- and three-compartment pharmacokinetic models for the parent and one- and two-compartment models for the metabolite were fitted to the data simultaneously to determine the best structural model. Elimination was assumed to occur from the central compartment as a first order process, and conversion of telapristone to CDB-4453 was considered to be an irreversible process.

Interindividual variability (IIV) of the pharmacokinetic parameters was modeled assuming a log-normal distribution, as follows:

\[ P_i = P_{pop} \cdot \exp(\eta_i) \]

Where \( P_i \) is the estimated parameter value for individual i, \( P_{pop} \) represents the typical population estimate for the parameter and \( \eta_i \) is the deviation of \( P_i \) from \( P_{pop} \). The \( \eta \) random effects were assumed to be independent and symmetrically distributed with a mean of zero and variance \( \sigma^2 \). The magnitude of IIV was expressed as coefficient of variation (\%CV).

Residual variability (RV) was modeled using a log error model as shown below:

\[ \ln C_{ij} = \ln C_{pred,ij} + \epsilon_{ij} \]

Where \( C_{ij} \) and \( C_{pred,ij} \) represent the jth observed and model predicted telapristone or CDB-4453 concentrations, respectively, for individual i. \( \epsilon_{ij} \) denotes the additive residual random error for individual i and observation j. The \( \epsilon \) random effects were assumed to be independent and symmetrically distributed with mean of zero and variance \( \sigma^2 \).

Covariate Analysis. Covariate analysis was performed in order to identify additional variables that might be able to explain the variability seen in the parameter estimates of the base model. Continuous covariates available for screening were age, weight, height, BMI, aspartate aminotransferase (AST) and alanine aminotransferase (ALT), and creatinine clearance (CrCl). Categorical covariates available were hepatic and renal impairment status. Covariates were initially screened for inclusion using visual inspection of the IIV versus covariates plots and generalized additive modeling using Xpose. The potential covariates were then tested for statistical significance using the stepwise forward addition, followed by stepwise backward elimination procedure (14). An MOFV change of 3.84 (corresponding to a significance level of 5% at 1 degree of freedom) was used as the cutoff to include a covariate in stepwise forward addition. For a covariate to remain in the model during backward elimination, a change in MOFV of at least 10.83 (corresponding to a significance level of 0.1% at 1 degree of freedom) was needed. An improvement in the precision of the parameter estimate (relative standard error), and reduction in intersubject and residual variability were also used to determine the importance of the covariate as a predictor.

Mixture Model. Comparison of total clearance of telapristone, via noncompartmental analysis, showed that generally, subjects could be separated into two clearance groups based on metabolic ratios (Fig. 2). The reason for the differences seen in clearance is unknown. Since inclusion of covariates did not account for the differences seen in total clearance of telapristone, a mixture model (using the SMIXTURE subroutine in NONMEM®) was employed in order to model two subpopulations for total clearance of telapristone. Because the NONMEM® output results only gives the overall probability of being in a particular group, individual objective value contributions were used to evaluate the probability of a subject being in each population. This was done using the method published by Carlson et al. (15). Where:

\[ OFV = \sum_{j=1}^{m} [2 \ln (\text{IL}_j)] \]

\[ \text{IL}_j = \sum_{k=1}^{m} \exp(-\text{OFV}_i,k/2) \times P_{pop,k} \]

\[ IP_k = \frac{\sum_{i=1}^{m} \text{IL}_i,k \times P_{pop,k}}{\sum_{i=1}^{m} \text{IL}_i,k \times P_{pop}} \]

the individual’s contribution to the objective function value (OFV) and estimated probability (P) evaluated for
all populations (pop. k) was used to calculate that individual’s likelihood (IL_i). These values were then used to calculate the individual’s probability of belonging to either population (IP_k).

**Model Evaluation.** Diagnostic plots used to assess model goodness-of-fit included observed concentrations versus population predictions (PRED), observed concentrations versus individual predictions (IPRED), conditional weighted residuals (CWRES) versus PRED, and CWRES versus time.

The standard errors of the final model parameter values were estimated using the non-parametric bootstrap approach in Perl-speaks-NONMEM® (PsN®) (16). For the final model, 500 bootstrap datasets were generated. The mean and 95% confidence intervals for the population pharmacokinetic parameters were calculated, and compared with the estimates from the final model. The bootstrap 95% confidence interval was calculated based on the percentile of the empirical distribution of the estimated parameters from the bootstrap runs.

**Model Prediction.** Visual predictive check was performed to evaluate the predictive ability of the final model. One thousand virtual observations at each sampling time point for the full dataset (including time points for BQL data) were simulated using the final model parameter estimates. The observed data were then plotted with the median and 90% prediction interval of the simulated data. Simulated concentrations below the quantification limit (BQL) were excluded, and not used in the calculation of the prediction intervals.

**Secondary Parameters.** Secondary parameters for each individual were calculated using maximum a posteriori Bayesian estimates for each individual. The equations below were used to calculate secondary parameters:

\[
\beta(h^{-1}) = \frac{1}{2} \left[ k_{24} + k_{42} + k_{el} - \sqrt{(k_{24} + k_{42} + k_{el})^2 - 4 \cdot k_{42} \cdot k_{el}} \right]
\]

\[\alpha(h^{-1}) = \frac{k_{el} \cdot k_{42}}{\beta} \quad \text{Where kel is the rate constant for loss from the central compartment}\]

\[t_{1/2d}(h) = \ln(2) / \beta \]

\[t_{1/2a}(h) = \ln(2) / \alpha \]

**Simulations.** Monte Carlo simulations were performed in order to reflect the expected range of variability in concentration values using the final model parameter estimates. Since telapristone is being developed for the chronic treatment of endometriosis and uterine fibroids, 500 simulations of telapristone and CDB-4453 were performed at steady state, 25 mg. Total concentration (nanomoles per liter) of telapristone and CDB-4453 were used in the analysis since CDB-4453 is an active metabolite with possible equipotent anti-progestational activity in vivo.

**RESULTS**

**Dataset**

Demographic and clinical data for subjects enrolled in both studies are provided in Table I. There were a total of 1,805 concentration measurements for studies ZP-005 and ZP-006. Ninety (5.0%) concentration measurements of telapristone and CDB-4453 were below the quantification limit (BQL).

**Population Pharmacokinetic Analysis**

**Model Development.** The final base model, shown in Fig. 2, was a two-compartment parent, one-compartment metabolite mixture model with first-order absorption without lag and first-order elimination. The model was parameterized in terms of absorption rate constant (Ka), oral clearance (CL/F), apparent volume of distribution of the central compartment (V2/F), apparent peripheral volume of distribution for the parent (V4/F), apparent inter-compartmental clearance (Q/F), and apparent peripheral volume of distribution for the central compartment (V3/F).

Since CDB-4453 was not administered alone and the true fraction of telapristone converted to CDB-4453 (Fmet) is
unknown, Fmet and V3/F are unidentifiable. Hence, the apparent volume of distribution of the metabolite (V3/F) was fixed to 1 L (17). Because V3/F (V_{app}) was fixed to 1 L, estimated Fmetest was then interpreted as the ratio of the fraction of telapristone converted to CDB-4453 to the distribution volume of CDB-4453 (Fmet/V_{app}).

**Covariate Analysis.** Renal and hepatic insufficiency, age, weight, height, BMI, ALT, and AST were tested for significance as covariates. Moderate renal impairment was found to be the only significant covariate and was incorporated using a proportional covariate model.

\[
K_a = \theta_1 \cdot (1 + \theta_2 \cdot \text{RENAL})
\]

Where \(\theta_1\) was the estimated value for absorption rate constant in healthy and mild renal-impaired subjects (\(\text{RENAL}=0\)) and \(\theta_2\) is the proportional constant for the presence of moderate renal impairment (\(\text{RENAL}=1\)). Inclusion of moderate renal impairment resulted in a 74% decrease in absorption rate constant when compared to healthy and mild renal impairment. Stochastic simulation of concentration time profiles for moderate renal impaired and healthy/mild renal impaired subjects showed that there was on average a lower total \(C_{max}\) in subjects with moderate renal impairment (Fig. 3). In addition, evaluation of post hoc estimates showed the moderate renal-impaired subjects had a mean \(K_a\) value of 0.36±0.20 h\(^{-1}\) and healthy/mild renal-impaired subjects had a mean \(K_a\) value of 1.27±0.36 h\(^{-1}\). It should be noted that there was only limited data available on mild (\(n=6\)) and moderate (\(n=6\)) renal-impaired subjects.

**Fig. 3.** Simulated concentration time profile for moderate renal impaired and healthy/mild renal impaired subjects at 25 mg
Mixture Model. Incorporation of a mixture model for total clearance of telapristone resulted in a 3.15-drop in the MOFV and IIV to decrease by 26.1% when compared to the non-mixture model. The population estimates for apparent oral clearance (CL/F) for the two populations were 11.6 and 3.37 L/h, respectively. Elimination half-life for the high- and low-clearance group was 12 and 35 h, respectively. The probability of being in the high-clearance group was 0.251, and the low-clearance group was 0.749. Assessment of the individual population probabilities showed that most subjects were separated into their respective clearance groups with very high probability.

Model Evaluation. The population pharmacokinetic parameters along with standard errors, interindividual variability, and residual variability are presented in Table II. Basic goodness-of-fit plots for telapristone and CDB-4453 are presented in Figs. 4 and 5. Overall, the model adequately described the data. There appeared to be some bias in the

| Parameter                  | Estimate | %RSE | Bootstrap estimate (standard error) | Bootstrap estimate (95% CI) |
|----------------------------|----------|------|-----------------------------------|-----------------------------|
| CL/F (L/h)—high clearance | 11.6     | 36.0 | 12.8 (4.63)                       | (6.50–24.28)                |
| CL/F (L/h)—low clearance  | 3.34     | 19.1 | 3.31 (0.637)                      | (2.42–4.67)                 |
| V2/F (L)                   | 37.4     | 10.2 | 37.5 (3.83)                       | (30.4–46)                   |
| V3/F (L)                   | 1        |      |                                   |                             |
| V4/F (L)                   | 120      | 12.4 | 123 (15.92)                       | (94.6–156)                  |
| Q/F (L/h)                  | 21.9     | 14.4 | 22.3 (2.88)                       | (16.1–29.0)                 |
| Ka (h⁻¹)                   | 1.26     | 7.21 | 1.27 (0.102)                      | (1.10–1.51)                 |
| Moderate renal impairment on Ka | -0.744  | 11.6 | -0.722 (0.174)                   | (-0.874–(-0.391))          |
| Fmetest (L⁻¹)              | 0.201    | 16.8 | 0.201 (0.032)                     | (0.137–0.267)               |
| CLM/F (L/h)                | 2.43     | 15.1 | 2.41 (0.343)                      | (1.73–3.14)                 |
| Probability                | 0.251    |      | 0.278 (0.131)                     | (0.038–0.534)               |
| IIV-CL/F                   | 0.200 (44.7) | 43.4 | 0.163 (0.073)                     | (0.046–0.318)               |
| IIV-V2/F                   | 0.443 (66.6) | 29.1 | 0.434 (0.142)                     | (0.207–0.739)               |
| IIV-V4/F                   | 0.426 (65.3) | 33.3 | 0.416 (0.144)                     | (0.165–0.736)               |
| IIV-Q/F                    | 0.457 (67.6) | 31.9 | 0.432 (0.145)                     | (0.160–0.731)               |
| IIV-Ka                     | 0.280 (52.9) | 45.7 | 0.263 (0.127)                     | (0.433–0.557)               |
| IIV-Fmetest                | 0.196 (44.3) | 22.5 | 0.189 (0.044)                     | (0.109–0.282)               |
| RV (log error)             |          |      |                                   |                             |
| Telapristone (nmol/L)      | 0.378    | 15.8 | 0.378 (0.061)                     | (0.267–0.516)               |
| CDB-4453 (nmol/L)          | 0.109    | 18.6 | 0.109 (0.020)                     | (0.073–0.156)               |

Fig. 4. Goodness-of-fit plots for telapristone (dashed lines are smoothing lines and solid lines are lines of identity)
predictions of telapristone concentrations at the earlier time points, but evaluation of other types of absorption profiles (first order with lag, mixed, parallel, and Weibull) lead to model instability, decrease in precision of parameter estimates, and/or unreasonable parameter estimates.

The parameter estimates with standard error and 95% confidence interval generated from the bootstrap method are presented in Table II. Ninety-nine percent of the runs converged successfully. The parameter estimates obtained from the final model were reasonably close to the ones obtained from the bootstrap analysis. All the parameter estimates from the final model were also contained within the 95% bootstrap intervals. For the visual predictive check, 7.95% of the telapristone observations and 6.00% of CDB-4453 observations were not contained within the 90% prediction interval (Fig. 6).

**Table III.** Steady State Pharmacokinetic Parameters (Using Total Concentration of Telapristone and CDB-4453) for 500 Simulated Subjects at 25 mg

| Parameter               | Median     | 5th and 95th Percentiles |
|-------------------------|------------|--------------------------|
| **High-clearance group**|            |                          |
| $C_{\text{max}}$ (nmol/L) | 1,273      | 538–2,750                |
| $T_{\text{max}}$ (h)       | 2.00       | 0.750–7.00               |
| $\text{AUC}_{\text{tot} 0\text{--}24}$ (nmol*h/L) | 9,141 | 5,507–16,288 |
| **Low-clearance group**   |            |                          |
| $C_{\text{max}}$ (nmol/L) | 2,308      | 1,099–5,266              |
| $T_{\text{max}}$ (h)       | 2.00       | 0.25–16                  |
| $\text{AUC}_{\text{tot} 0\text{--}24}$ (nmol*h/L) | 22,458 | 11,238–46,691 |
Simulations. For steady-state simulation of telapristone and CDB-4453, total C\text{\textsubscript{max}} and AUC\textsubscript{0–24} were higher in the low-clearance group (Table III).

DISCUSSION

A two-compartment (parent) one-compartment (metabolite) mixture model with first-order absorption and elimination adequately described the pharmacokinetics of telapristone and CDB-4453. There appeared to be some bias in the predictions of telapristone concentrations at the earlier time points such that concentrations were underpredicted. Inspection of individual concentration-time profiles (observed, individual-predicted, and population-predicted values) did not show any systematic biases; however, there were some subjects that appeared to exhibit a lag time while others did not.

Since F\text{\textsubscript{met}} (F\text{\textsubscript{met/V\text{\textsubscript{app}}}}) was estimated to be 0.20 L \text{\textsuperscript{-1}} and F\text{\textsubscript{met}} can be no greater than 1; the apparent volume of distribution of the metabolite (V\text{\textsubscript{3/F}}) cannot be greater than 5 L. Therefore, it appears that the central volume of distribution of the metabolite is smaller than that of the parent. The removal of methyl group from a parent compound could lead to a more polar metabolite. Therefore, the smaller volume of distribution of CDB-4453 could be due to increased hydrophilicity as a result of the removal of the methyl group.

The only significant covariate found in this analysis was moderate renal impairment on absorption rate constant. The effect of renal impairment on absorption rate constant has been reported previously in literature (18–22). Changes in the gastrointestinal tract as a result of renal insufficiency could result in delayed gastric emptying. This, in turn, can lead to a decrease in the rate of absorption of the drug, and hence increase time to reach maximal concentration (T\text{\textsubscript{max}}). Although delayed gastric emptying can lead to an increase in T\text{\textsubscript{max}} it generally does not affect the extent of absorption (22–24). A delay in T\text{\textsubscript{max}} might be problematic in drugs that are administered acutely, as the onset of a response might be delayed. However, for chronic treatment, as is the case of telapristone, the effect of a decrease in absorption rate constant might be clinically irrelevant. Since creatinine clearance did not meet the inclusion criteria to be tested as a covariate, it is believed that the effect of moderate renal impairment is not directly related to the disease state (e.g., changes in glomerular filtration rate). Rather, it might be a reflection of other pathophysiological states precluding or resulting from chronic renal impairment. The lack of significance of hepatic impairment as a covariate is important since it is believed that the majority of telapristone is primarily metabolized in the liver. This suggests that dosing adjustment may not be required on a pharmacokinetic basis in patients with mild or moderate hepatic impairment. Although the ranges of values for the covariates tested were wide, the sample size was very small. Therefore, a larger sample size might be required to detect additional covariate effects.

The detection of two subpopulations for total clearance might reflect the polymorphic metabolism of telapristone to CDB-4453. The percentage of the population in the low- and high-clearance groups was 75% and 25%, respectively. Since it is proposed that telapristone utilizes the same metabolic pathways as mifepristone (10,11), the cytochrome P450 (CYP) that would be responsible for production of the major metabolites of telapristone is CYP3A4 and 3A5. Even though there is known pharmacokinetic interindividual variability in drugs that are substrates for CYP3A4, there are no known polymorphisms that result in phenotypic differences in drug metabolism (25,26). On the other hand, CYP3A5 is known to be polymorphic and have overlapping substrates with CYP3A4 (26–28). Studies show that functioning CYP3A5 is present in 10–40% of the Caucasian population, approximately 50% of African Americans, and about 33% of Asians (26,29,30). Individuals that express CYP3A5 also express CYP3A4 and it is believed that the presence of both of these CYPs lead to an increase in clearance of drugs that are substrates for both enzymes (26). Based on the percentages of the study population in the high- and low-clearance groups, it is hypothesized that, like mifepristone, telapristone is both a substrate for CYP3A4 and CYP3A5. Additional studies are needed to evaluate telapristone metabolic pathways.

Because telapristone acetate is being developed for the chronic treatment of symptoms associated with leiomyoma and endometriosis, steady-state simulations are important for evaluating the pharmacokinetics of telapristone and CDB-4453. The resulting higher total C\text{\textsubscript{max}} and AUC\textsubscript{0–24} in the low-clearance group compared to the high-clearance group at steady state could be the result of the longer half-life of telapristone seen in the low-clearance group. A longer half-life could lead to greater accumulation of the drug. Therefore, the presence of two groups for total clearance may play an important role in the efficacy/toxicity of telapristone and CDB-4453 with long-term treatment.

CONCLUSIONS

In conclusion, a parent-metabolite mixture model adequately described the pharmacokinetics of telapristone and CDB-4453 after oral administration. Inclusion of a mixture model resulted in detection of two subpopulations for the total clearance of telapristone. Inclusion of renal impairment as a covariate resulted in a decrease in the absorption rate constant of telapristone. It is believed that the significance of moderate renal impairment on absorption rate constant was not directly attributed to renal impairment (in terms of decrease glomerular filtration rate/changes in creatinine clearance). Rather, the changes in absorption rate constant, was attributed to some other disease precluding/resulting in renal impairment or other pathophysiologic states induced by renal impairment (e.g., delayed gastric emptying as a result of diabetes).

CONFLICT OF INTEREST

Joseph Podolski, Alan Kirsch, and Ronald Wiehle are employees of Repros Therapeutics.

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