Organic food use, meat intake, and prevalence of gestational diabetes: KOALA birth cohort study

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Abstract: Purpose: To evaluate whether consumption of organic food and reduced intake of meat products in pregnancy are associated with lower prevalence of gestational diabetes (GD). Methods: Women participating in the KOALA Birth Cohort Study with valid informed consent, a singleton pregnancy and information on their food intake were considered in this cross-sectional analysis. Participants with and without GD were compared with each other in terms of dietary characteristics (n = 37 and n = 2766, respectively). Multivariable logistic regression (LR) was used to adjust for relevant covariates. Results: Organic food consumption tended to be lower, although not significantly, in women with GD compared to women without GD, whereas consumption of meat was positively associated with GD prevalence. LR modelling showed that GD was significantly associated with higher consumption of meat and, in addition, also of cheese, after adjustment for other relevant covariates. GD was associated with some indicators of animal product intake, namely dietary animal to plant protein ratio and maternal plasma arachidonic acid (for the latter, data available for n = 16 and n = 1304, respectively). Food patterns of participants with GD were characterised by more meat products and less vegetarian products. Conclusions: Due to the low number of participants with GD, results have to be interpreted cautiously. Consumption of organic food during pregnancy does not seem to be markedly associated with a lower GD prevalence; lower intake of meat and cheese, irrespective of its origin (organic or conventional), does. The latter supports previous studies suggesting a causal association between consumption of animal products and GD. Keywords: Diet composition; Food patterns; Gestational diabetes; Meat consumption; Organic food.

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ORGINAL CONTRIBUTION

Organic food use, meat intake, and prevalence of gestational diabetes: KOALA birth cohort study

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Abstract

Purpose To evaluate whether consumption of organic food and reduced intake of meat products in pregnancy are associated with lower prevalence of gestational diabetes (GD).

Methods Women participating in the KOALA Birth Cohort Study with valid informed consent, a singleton pregnancy and information on their food intake were considered in this cross-sectional analysis. Participants with and without GD were compared with each other in terms of dietary characteristics (n = 37 and n = 2766, respectively). Multivariable logistic regression (LR) was used to adjust for relevant covariates.

Results Organic food consumption tended to be lower, although not significantly, in women with GD compared to women without GD, whereas consumption of meat was positively associated with GD prevalence. LR modelling showed that GD was significantly associated with higher consumption of meat and, in addition, also of cheese, after adjustment for other relevant covariates. GD was associated with some indicators of animal product intake, namely dietary animal to plant protein ratio and maternal plasma arachidonic acid (for the latter, data available for n = 16 and n = 1304, respectively). Food patterns of participants with GD were characterised by more meat products and less vegetarian products.

Conclusions Due to the low number of participants with GD, results have to be interpreted cautiously. Consumption of organic food during pregnancy does not seem to be markedly associated with a lower GD prevalence; lower intake of meat and cheese, irrespective of its origin (organic or conventional), does. The latter supports previous studies suggesting a causal association between consumption of animal products and GD.

Keywords Gestational diabetes · Organic food · Meat consumption · Diet composition · Food patterns

Abbreviations

BMI Body mass index
FFQ Food frequency questionnaire
GD Gestational diabetes
KOALA ‘Kind, Ouders en Gezondheid: Aandacht voor Leefstijl en Aanleg’ (child, parents and health, addressing lifestyle and constitution)
LR Logistic regression
NEVO-1 Dutch food consumption table aggregate level 1
PCA Principal component analysis

Introduction

Gestational diabetes (i.e. diabetes with first diagnosis or onset during pregnancy, GD) accounts for most cases of diabetes during pregnancy and is one of the most common pregnancy-related disorders. GD-associated perinatal complications include macrosomia, hypoglycaemia, respiratory distress syndrome, polycythaemia, hyperbilirubinemia, cardiomyopathy and infant death. In the last years, implications of in utero exposure to maternal GD for offspring long-term health have been unveiled. Research shows that GD can be associated with future adiposity, adverse cardiometabolic outcomes and reduced cognitive ability [1–3]. GD prevalence has been increasing in the last decade and at present,
it occurs in approximately seven percent of all pregnancies [4]. This increase accompanies the worldwide rise in type 2 diabetes prevalence, which is likely to be related to a rapid replacement of traditional diets by dietary patterns with high proportions of animal products, processed industrial foods, refined sugars, saturated fats and oils [5–7]. There are some indications that dietary patterns can influence GD development as well. Dietary patterns characterised by high amounts of whole grains, fruits, and vegetables seem to reduce the GD risk, and Western dietary patterns as well as those characterised among others by high amounts of (red and/or processed) meat are associated with higher prevalence [8–10]. Currently, high meat consumption deserves special attention, because evidence for a direct negative impact on human health is accumulating [11–15] and meat production practices are unsustainable [16]. Discussions about how much meat should be consumed and how to produce meat products in a sustainable way are needed [17].

Production of organic foods relies on ecological processes, biodiversity, and cycles adapted to local conditions. At the same time, agricultural inputs with unknown or adverse effects are avoided. These include genetically modified seeds, synthetic fertilisers and pesticides, preventive veterinary drugs, and most preservatives, flavour enhancers and other additives, and irradiation during processing [16]. There are still diverging opinions about sustainability of organic agriculture [18, 19] and the impact of consumption of organic food on health (reviewed in [20]). A few studies suggest the presence of higher amounts of health-promoting substances, such as favourable fatty acids and vitamins, in organic food (see [20] and references therein). The lower use of synthetic fertilisers and pesticides, veterinary drugs with preventive objectives, and preservatives and additives is per se promising. Recent work reveals that—likely due to the low levels of synthetic pesticides—consumption of organic food is inversely associated with the risk of type 2 diabetes [21].

Several studies were initiated to investigate health-related effects of organic food consumption during pregnancy, a period in which health expectations are particularly relevant, since they can long-term affect both mothers’ and their offspring’s health. Results suggest favourable impacts of organic food consumption on the prevalence of pre eclampsia [22] and child hypospadias [23], as well as on the later development of child atopic sensitisation, allergies and eczemas [24–26]. Moreover, our previous work with data from the Dutch KOALA-study revealed that organic food consumption is associated with lower prevalence of maternal overweight and adiposity during pregnancy [27], which suggests a beneficial effect of food of organic production on GD development. Although the results obtained with the subgroup of KOALA-participants for which additional health-related biomarkers are available were in line with this hypothesis [28], the sample size was not big enough for its definitive verification.

Analyses of European cohort studies from France [29], Germany [30], Norway [31], and the Netherlands [27] on organic food consumption are unanimous in showing that intake of organic food goes hand in hand with lower consumption of meat and meat products and with specific food patterns. Therefore, it is highly challenging to try to disentangle whether associations between organic food consumption and health-related outcomes can be truly attributed to the organic food by itself, or otherwise to the dietary composition or overall food patterns that are associated with the consumption of organic food. In our previous study, we showed that part of—but not all—the observed associations between organic food consumption and health-related biomarkers could be explained by the food patterns accompanying the consumption of organic food [28]. Following this idea, and using data from the whole KOALA-cohort, we now aimed at investigating the associations between (i) the consumption of organic food, (ii) the consumption of meat and meat products, (iii) dietary composition, (iv) dietary indicators and an independent blood plasma indicator of animal product intake, and (v) food patterns, and the development of GD.

Methods

Study design and participants

Recruitment of healthy, pregnant women took place from October 2000 onwards, at midwives’ practices in the southern part of the Netherlands (n = 2434). They were participating in an ongoing prospective cohort study on Pregnancy-Related Pelvic Girdle Pain. This study did not use lifestyle-related inclusion criteria and the participants showed a good comparability with the general population [32]. This group is referred to as the general population recruitment group. To enrich the KOALA-study with participants adhering to an alternative lifestyle, 491 pregnant women were recruited from October 2001 onwards through various specific channels (alternative recruitment group). Those specific channels included organic food shops, anthroposophic general practitioners, midwives, and under-five clinics, Rudolf Steiner schools and advertisements in magazines on organic food and other relevant magazines. Recruitment of both groups continued to end of 2002. All women participating in the KOALA-study were enrolled between 14 and 18 weeks of gestation and received detailed questionnaires on sociodemographic and health-related characteristics at the time of recruitment as well as at 30 and 34 weeks of pregnancy (n = 2993). The questionnaire issued at 34 weeks of pregnancy dealt with lifestyle and dietary...
habits during the previous month, including items on organic food purchase and consumption, and a food frequency questionnaire (FFQ).

The population flow chart showing the stepwise recruitment of the study sample can be found in the Supplementary information (Supplementary material 1). Among the 2993 enrolled pregnant women with valid informed consent, 37 had a twin pregnancy, leaving 2956 singleton pregnancies. From these, complete dietary data (FFQ) at week 34 were available for 2821 women; 135 women with no or incomplete dietary data had to be excluded. Also excluded were 14 women whose children turned out to have Down’s syndrome (n = 7), or any other congenital disorder (n = 7). This means that 2807 women with a singleton pregnancy, available FFQ data and a child without a diagnosed congenital disease were left for further processing. Of the considered 2807 women, 2333 have been recruited from the conventional lifestyle group and 474 from the alternative lifestyle recruitment group.

The primary outcome was gestational diabetes (GD). GD was defined as self-reported gestational diabetes at week 34 (i.e. reported diabetes and absence of diabetes before pregnancy) and/or diagnosis of GD mentioned in the midwife reports. Information on GD status (both self-reported and in midwife reports) was missing in 4 women and 3 of the women without GD had been diagnosed with diabetes mellitus before the current pregnancy and were therefore not at risk of developing GD. This gave rise to exclude another 7 women from the multivariate LR, which was therefore performed with data from 2800 pregnant women.

In all women recruited from January 2002 onwards, blood samples were collected at 34–36 weeks of pregnancy and plasma fatty acids were measured by gas chromatography as described in detail elsewhere [28]. In the present analysis, plasma value of arachidonic fatty acid was included as an independent indicator of animal product intake. These data were available for 1320 (47.1%) of the 2800 women in the analysis.

### Dietary measurements

From the week 34 questionnaire and as described in detail elsewhere [27], answers to the questions dealing with the following topics were used:

1. **Origin (conventional or organic) of products from the following seven food groups consumed during the previous month:** meat; eggs; vegetables; fruit; milk and milk products; bread; and dried foods (dried legumes and cereals). For each of these food groups, we asked participants to report the percentage of purchased products from organic origin: less than 50%, between 50 and 90%, or more than 90%.

2. **Habitual food intake.** Intake of food products was assessed based on a 198-item FFQ that is an extended version of a previously existing, validated FFQ [33]. For each of the inquired food products the frequency of use (nine categories: no use, 1–3 days per month, and 1, 2, 3, 4, 5, 6 and 7 days per week) during the previous month had to be recalled and noted down. In addition, the average amount per consumption day was asked for. Depending on the food product this could be done in grams (e.g., grams of meat), natural units (e.g., number of apples, slices of bread), or household units (e.g., glasses of milk, spoons of rice). All food products were classified according to the Dutch Food Consumption Table aggregate level 1 (NEVO-1 food groups). Food products within the same food level 1-group were taken together to obtain the daily intake of each group per participant (in g/day; see [27] for further details). Relative frequency of use of food subtypes was asked for part of the foods.

3. **Energy intake, micronutrients and macronutrients.** The average daily dietary intake of energy (kJ), micronutrients (mg, µg) and macronutrients (g) was calculated by combining for each FFQ item the frequency and amount data with standard portion sizes (grams) and with nutrient composition per 100 g of food as documented in the 2010 online version of the Dutch Food Composition Table for about 170 nutrients (Voedingscentrum 2010); as in previous work [27], supplements were not considered in this calculation. In the present analysis, energy intake and the data required to calculate dietary ratios haem to non-haem iron and animal to plant protein (indicators of animal product intake) were considered.

### Definition of subgroups of organic food intake and meat intake

As in our previous work [27], the use of organic food was defined based on the questions concerning the origin of the seven food groups mentioned above (meat; eggs; vegetables; fruit; milk and milk products; bread; and dried foods) and on the % organic origin of these seven food groups taken together. The following subgroups were defined according to the consumption of organic food: “ < 50% organic”, if some food groups were of organic origin, but not all were reported as being more than 50% organic; “ > 50% organic”, if in all consumed food groups at least 50% was of organic origin [27]. While doing so, only the consumed food groups were considered (a few participants reported not consuming a specific food group, e.g. meat). The remaining group of participants who used all consumed food groups mostly of conventional origin was considered as the “conventional” group.
Regarding meat intake, participants were categorised in quartiles according to their consumption of meat, processed meat and poultry (food group 23 from the 2010 NEVO-1 classification).

**Statistical methods**

Differences in participant characteristics between the groups without (reference) and with GD were evaluated by analysis of variance (ANOVA) and chi-squared test for continuous and categorical variables, respectively.

To identify NEVO-1 food groups associated with GD, multivariable logistic regression was used with GD as the outcome variable, and a stepwise backward elimination of NEVO-1 food groups, using an F-to-enter (PIN) of 0.05 and a F-to-remove (POUT) of 0.10 as the criteria to include or remove a variable in each consecutive step.

Multivariable logistic regression analysis was used to estimate the strength of the association of organic food use and meat intake, with GD status as the outcome, adjusting in first instance only for maternal constitutional factors (maternal age, gravidity). To adjust for possible confounding by (other) food groups, we first identified NEVO-1 food groups associated with GD (see above). These food groups were added to the initial models of organic food and meat intake, respectively. As a final step, other potential confounding covariates were added to derive the final, fully adjusted model, including: maternal age, gravidity, organic food consumption, maternal smoking in pregnancy, living area, season of filling out the questionnaire, and recruitment group and period (to adjust for secular trend: conventional group recruited from October 2000 to October 2001; conventional group recruited from October 2001 to end of 2002; alternative group recruited from October 2001 to end of 2002).

To evaluate possible influences of energy intake and of pre-pregnancy BMI, we used sensitivity analyses by comparing the final model with and without total energy intake and with and without pre-pregnancy BMI. Additional logistic regression analyses were used to estimate odds ratios for the association between selected indicators of animal product intake (dietary ratios haem to non-haem iron and animal to plant protein and plasma value of arachidonic fatty acid) and GD, with and without adjustment for the covariates shown in table footnotes.

Multivariable analysis was performed to calculate values that could be used to substitute missing values for any of the food items \((n = 15, \text{ one food item missing per participant})\). Having imputed the calculated values, Principal Components Analysis (PCA) of all food items was performed to further characterise food patterns. After setting the threshold on 0.15, the varimax rotation with Kaiser-normalisation led to the following results. The Kaiser–Meyer–Olkin measure verified the sample adequacy for analysis \((KMO = 0.698)\) and the Bartlett’s test of sphericity \(X^2(225/578) = 97.741.962, p = 0.000\), indicated that correlations between items were sufficiently high for PCA.

Data were analysed with IBM® SPSS® Statistics Version 19, except the logistic regression analyses, for which Version 25 was used. Statistical significance was set at \(p \leq 0.05\).

**Results**

**Main participant characteristics**

Women with and without GD were comparable in terms of recruitment characteristics, maternal age, level of education and living area (Table 1). In comparison with women without GD (reference group), women with GD had a slightly shorter pregnancy duration (mean ± standard deviation of 39.0 ± 1.4 weeks versus 39.5 ± 1.5 weeks) and higher average gravidity (2.3 ± 1.7 versus 1.9 ± 1.0). Clear-cut differences were observed with regard to mothers’ BMI before pregnancy \((27.3 ± 5.8 \text{ versus } 23.6 ± 3.9)\). Energy intake was similar in the two groups.

**Consumption of organic food and meat intake in relation to gestational diabetes status**

When the consumption of organic products in participants with and without GD was compared, a trend of lower organic food consumption in GD women was observed, which, however, was not significant (see Table 2). Further adjustments for NEVO-1 food groups and other covariables did not essentially alter these results (not shown).

The proportion of women with GD increased successively across the quartiles of meat intake (NEVO-1 group 23), with a consistent dose–response relationship towards higher GD risk associated with higher meat intake (test-for-trend over the quartiles \(p = 0.006\), initial adjusted model, Table 2). Among the participants in the fourth quartile, with the highest meat consumption \((136–367 \text{ g/d})\), the prevalence of GD was of 16 out of 700, i.e. 2.3%, whereas the prevalence was only 3 out of 700, i.e. 0.4%, among the participants in the first quartile.

**Dietary composition versus gestational diabetes**

Following an explorative approach, we investigated the association between the intake of products from various food groups and GD status. Since intakes from various NEVO-1 food groups are often inter-related, we initially considered to include all food groups—not only meat—in the multivariable analysis. A stepwise elimination resulted in the retention of six food groups in the final model. NEVO-1 food groups positively associated with GD were:
meat/meat products/poultry (group 23), cheese (group 11) and drinks (group 2), whereas intake of sugar/sweets/sweet sauces (group 20), cereals/cereal products (group 8), and pastry and biscuits (group 7) were negatively associated with GD status (Table 3, unadjusted). These results were in most cases confirmed by a comparison of the diet composition in terms of NEVO-1 groups between participants with and without GD (Supplementary material 2), that also revealed higher consumption of eggs (group 5, borderline significance). In the final, fully adjusted model, both meat and cheese intake remained strongly associated with GD risk. The odds for GD was 2.67 (95%CI 1.14, 6.26) higher with each 100 g/day increase of meat intake, and 5.68 (1.54, 20.9) for each 100 g/day increase of cheese intake, independent from each other and from the intake of the other food groups (Table 3, adjusted). Total energy intake was not associated with GD (Table 3, adjusted), and the results for meat and cheese did not change when energy intake was eliminated from this model (results not shown). Finally, adding the pre-pregnancy BMI to this model, only slightly attenuated the odds for GD in the case of meat (from 2.67 to 2.41; p values 0.02 and 0.04) and even slightly increased them in the case of cheese (from 5.68 to 5.95; p value 0.01 in both cases).

Dietary and independent blood plasma indicators of animal product intake

We reasoned that the participants who consumed higher amounts of meat/meat products/poultry and cheese (NEVO-1 groups 23 and 11) would exhibit higher values of three indicators of animal product intake: dietary ratios haem to non-haem iron and animal to plant protein, and maternal plasma arachidonic acid levels. The correlations between these indicators and the consumption of products from NEVO-1 groups 23 and 11 were investigated by calculating Pearson correlation coefficients (Supplementary material 3). The correlations were statistically significant for the haem to non-haem iron ratio, and the动物 to non-animal protein ratio (correlation coefficients 0.62 and 0.53, respectively). The maternal plasma arachidonic acid concentration showed a weak but significant association with both meat and cheese intake (correlation coefficients 0.14 and 0.08). The haem to non-haem iron ratio and the animal to non-animal protein

| Characteristic | With gestational diabetes (n=37) | Without gestational diabetes (n=2766) | p value |
|---------------|---------------------------------|-------------------------------------|---------|
|               | n or (mis) mean ± SD or %       | n or (mis) mean ± SD or %           |         |
| Recruitment group and period |                               |                                     | 0.15*   |
| General recruitment group/early period | 13 35.1 | 1227 44.4 |         |
| Alternative recruitment group/late period | 4 10.8 | 470 17.0 |         |
| General recruitment group/late period | 20 54.1 | 1069 38.6 |         |
| Age of the mother at delivery (years) | 37 32.24 ± 3.9 | 2766 31.5 ± 4.1 | 0.24 |
| Pregnancy duration (weeks) | 36 39.0 ± 1.4 | 2754 39.5 ± 1.5 | 0.04 |
| Missings | 1 | 12 | 0.15 |
| Highest education |                               |                                     |         |
| Lower | 6 16.2 | 287 10.4 | 0.15 |
| Middle | 20 54.1 | 1050 38.0 |         |
| Higher | 7 18.9 | 929 33.6 |         |
| Academic | 4 10.8 | 374 13.5 |         |
| Others | 0 0.00 | 126 4.6 |         |
| Gravidity | 37 2.3 ± 1.7 | 2766 1.9 ± 1.0 | 0.02 |
| Living area (province) |                               |                                     | 0.43   |
| South (Limburg) | 14 37.8 | 1128 40.8 |         |
| Other | 23 62.2 | 1638 59.2 |         |
| Mothers’ BMI before pregnancy |                               |                                     | < 0.001 |
| Missings | 35 27.3 ± 5.8 | 2700 23.6 ± 3.9 |         |
| Energy intake at gestational week 34 (kJ/day) | 37 10,055 ± 1931 | 2766 10,591 ± 2586 | 0.21 |

Statistically significant p-values are written in bold
*p = 0.095 if only the late period subgroups are considered
ratio were higher in the participants with GD (Supplementary material 4); both dietary indicators were positively associated with GD risk (although the first one lost the borderline statistical significance when adjusting for maternal age and gravidity; Table 4). Maternal plasma arachidonic acid was strongly associated with GD risk ($p = 0.01$, both unadjusted and adjusted, Table 4).

**Food patterns versus gestational diabetes**

Finally, we wanted to gain a more real insight into the everyday diet composition of the participants with and without GD. In the PCA performed to characterise food patterns, seven components were found (with eigenvalues over Kaiser’s criterion higher than 2; supported by the scree plot, not shown) that together explained 13.2% of the food items variance. Based on the underlying constituents, we labelled these main components as follows: “lacto-ovo-vegetarian”, “fast food”, “raw vegetables salad”, “fish”, “cooked vegetables”, “Italian-like kitchen and sweet”, and “meat” (see corresponding food items described in the order of decreasing absolute factor loadings in Supplementary material 5, both for positive and for negative loadings). The scores for “meat” were higher in women with GD than in those without, whereas scores for “lacto-ovo-vegetarian” and “Italian-like kitchen and sweet” were lower (Supplementary material 6).

| Characteristic | OR (95% CI) unadjusted | $p$ value | OR (95% CI) adjusted$^a$ | $p$ value |
|----------------|------------------------|----------|--------------------------|----------|
| Group 23, meat/meat products/poultry | 2.19 (1.2, 4.00) | 0.01 | 2.67 (1.14, 6.26) | **0.02** |
| Group 11, cheese | 3.87 (1.26, 11.8) | 0.02 | 5.68 (1.54, 20.9) | **0.01** |
| Group 2, drinks | 1.06 (1.01, 1.10) | 0.01 | 1.05 (0.99, 1.12) | 0.12 |
| Group 20, sugar/sweets/sweet sauces | 0.47 (0.24, 0.91) | 0.03 | 0.69 (0.31, 1.55) | 0.37 |
| Group 8, cereals/cereal products | 0.31 (0.11, 0.90) | 0.03 | 0.33 (0.10, 1.13) | 0.08 |
| Group 7, pastry and biscuits | 0.24 (0.05, 1.14) | 0.07 | 0.31 (0.06, 1.63) | 0.17 |
| Energy intake (100 kJ/day) | – | – | 0.99 (0.96, 1.01) | 0.24 |

Participants with diabetes before pregnancy were excluded from the logistic regression models

Statistically significant $p$-values are written in bold

$^a$ Covariates included in the model: energy intake, mother’s age at delivery, gravidity, recruitment group, organic food consumption, smoking, living area and season of filling out the questionnaire

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**Table 2** Associations between consumption of organic food and gestational diabetes status and between meat intake and gestational diabetes status, expressed as number of cases relative to total (and %; $n = 2803$) and odds ratio obtained in logistic regression models (and 95% CI, $n = 2800$)

| Characteristic | No. of cases with GD/total (%) | OR (95% CI) unadjusted | OR (95% CI) adjusted$^b$ |
|----------------|-------------------------------|------------------------|--------------------------|
| Organic food consumption | | | |
| Conventional | 25/1728 (1.5%) | 1.00 (reference) | 1.00 (reference) |
| $<50\%$ | 10/851 (1.2%) | 0.81 (0.39, 1.69) | 0.71 (0.32, 1.54) |
| $>50\%$ | 2/224 (0.9%) | 0.61 (0.14, 2.61) | 0.53 (0.12, 2.30) |
| Test-for-trend | | $p = 0.42$ | $p = 0.26$ |
| Meat consumption (quartiles of NEVO-1 group 23) | | | |
| Q1 0–74 g/d | 3/700 (0.4%) | 1.00 (reference) | 1.00 (reference) |
| Q2 74–106 g/d | 8/701 (1.1%) | 2.69 (0.71, 10.18) | 2.74 (0.72, 10.41) |
| Q3 106–136 g/d | 10/702 (1.4%) | 3.36 (0.92, 12.25) | 3.47 (0.95, 12.72) |
| Q4 136–368 g/d | 16/700 (2.3%) | 5.44 (1.58, 18.76) | 5.22 (1.48, 18.37) |
| Test-for-trend | | $p = 0.003$ | $p = 0.006$ |

Statistically significant $p$-values are written in bold

$^b$ Participants with diabetes before pregnancy were excluded from the logistic regression models

$^a$ Participants with diabetes before pregnancy were excluded from the logistic regression models

$^b$ Covariates included in the model: mother’s age at delivery and gravidity
A major limitation of this analysis is the modest number of independently from meat intake. Moreover, GD is associated with higher consumption of meat, even after confirmed that the risk of GD is statistically significantly in which the most relevant food groups were included, additional LR models, a direct dependency between consumption of meat products and GD prevalence was found. Additional LR models, in which the most relevant food groups were included, confirmed that the risk of GD is statistically significantly associated with higher consumption of meat, even after adjustment for confounders. In addition, we found that cheese consumption is associated with GD status as well, independently from meat intake. Moreover, GD is associated with dietary (haem to non-haem iron ratio, animal to plant protein ratio) and plasma (arachidonic acid levels) indicators of animal product intake. Finally, these findings were corroborated with a PCA analysis, which showed that food patterns of GD participants are characterised by higher meat product intake and lower vegetarian product intake.

A major strength of the present work is the combination of various approaches to investigate associations between diet characteristics and GD. A high contrast of participants’ diet composition, the very detailed FFQ used in the KOALA study, and the availability of data on macronutrients and micronutrients and (for some participants) on blood plasma biomarkers enabled that combination. Based on the nutrients’ data, two dietary indicators of animal product intake could be calculated. Among the plasma biomarkers, the semi-essential fatty acid arachidonic acid has the advantages that its main dietary source are animal products [34, 35], at the same time that, in humans, the ingested amounts of the precursor linoleic acid hardly affect its plasma levels [36]. A major limitation of this analysis is the modest number of women with GD (n = 37 for diet characteristics) and hence a limited statistical power to prove or rule out a modest effect of organic food use, although for meat intake the associations appeared strong enough to reach statistical significance even with the limited number of cases. In most cases of GD (20 out of 37), the diagnosis was supported by midwife reports, however in 17 cases, the diagnosis was self-reported only, which can be seen as a study limitation. Another limitation is that the pregnant women filled out the FFQ at pregnancy week 34, i.e. when a GD diagnosis is usually known. Since dietary changes are one of the first measures taken in GD treatment, it is likely that women with GD had already adapted their diet by the time of completing the FFQ. This might explain why—in the unadjusted MLR model—the GD participants seemed to consume lower amounts of sugar/sweets/sweet sauces, cereals, and pastry and biscuits; also, the somewhat higher amounts of drinks ingested might be related to GD-associated thirst. Importantly, increasing meat and cheese consumption is not pursued by GD treatment, so it is not likely that our results on meat and cheese intake are due to reverse causation.

Due to the low number of participants with GD, our results have to be interpreted cautiously. Nevertheless, they suggest that GD prevalence could be lower if women would reduce their meat and cheese intake. If the effect of meat consumption on GD risk would be causal, the effect size in this study (OR 2.83 per 100 g/day) could be interpreted as a two to three times lower risk of GD with lowering meat intake from 150 g/day to 50 g/day (i.e. from approximately the population mean plus 1 SD to the population mean minus 1 SD; compare with Supplementary material 2); or as a five times lower risk for pregnant women in the lowest quartile of meat consumption compared to women in the highest quartile (OR 5.22). For cheese intake, the effect size was found to be in the same order of magnitude (OR 5.14 per 100 g extra cheese per day). Our results on meat and cheese are in line with studies showing that Mediterranean Diet, that comprises few animal products, reduces GD incidence [37, 38]. Furthermore, they are supported by previous work.

Discussion

Our results show that, while consumption of organic food does not seem to markedly affect GD prevalence, higher intake of meat and meat products is associated with a higher prevalence. Based on logistic regression modelling, a direct dependency between consumption of meat products and GD prevalence was found. Additional LR models, in which the most relevant food groups were included, confirmed that the risk of GD is statistically significantly associated with higher consumption of meat, even after adjustment for confounders. In addition, we found that cheese consumption is associated with GD status as well, independently from meat intake. Moreover, GD is associated with dietary (haem to non-haem iron ratio, animal to plant protein ratio) and plasma (arachidonic acid levels) indicators of animal product intake. Finally, these findings were corroborated with a PCA analysis, which showed that food patterns of GD participants are characterised by higher meat product intake and lower vegetarian product intake.

A major strength of the present work is the combination of various approaches to investigate associations between diet characteristics and GD. A high contrast of participants’ diet composition, the very detailed FFQ used in the KOALA study, and the availability of data on macronutrients and micronutrients and (for some participants) on blood plasma biomarkers enabled that combination. Based on the nutrients’ data, two dietary indicators of animal product intake could be calculated. Among the plasma biomarkers, the semi-essential fatty acid arachidonic acid has the advantages that its main dietary source are animal products [34, 35], at the same time that, in humans, the ingested amounts of the precursor linoleic acid hardly affect its plasma levels [36]. A major limitation of this analysis is the modest number of women with GD (n = 37 for diet characteristics) and hence a limited statistical power to prove or rule out a modest effect of organic food use, although for meat intake the associations appeared strong enough to reach statistical significance even with the limited number of cases. In most cases of GD (20 out of 37), the diagnosis was supported by midwife reports, however in 17 cases, the diagnosis was self-reported only, which can be seen as a study limitation. Another limitation is that the pregnant women filled out the FFQ at pregnancy week 34, i.e. when a GD diagnosis is usually known. Since dietary changes are one of the first measures taken in GD treatment, it is likely that women with GD had already adapted their diet by the time of completing the FFQ. This might explain why—in the unadjusted MLR model—the GD participants seemed to consume lower amounts of sugar/sweets/sweet sauces, cereals, and pastry and biscuits; also, the somewhat higher amounts of drinks ingested might be related to GD-associated thirst. Importantly, increasing meat and cheese consumption is not pursued by GD treatment, so it is not likely that our results on meat and cheese intake are due to reverse causation.

Due to the low number of participants with GD, our results have to be interpreted cautiously. Nevertheless, they suggest that GD prevalence could be lower if women would reduce their meat and cheese intake. If the effect of meat consumption on GD risk would be causal, the effect size in this study (OR 2.83 per 100 g/day) could be interpreted as a two to three times lower risk of GD with lowering meat intake from 150 g/day to 50 g/day (i.e. from approximately the population mean plus 1 SD to the population mean minus 1 SD; compare with Supplementary material 2); or as a five times lower risk for pregnant women in the lowest quartile of meat consumption compared to women in the highest quartile (OR 5.22). For cheese intake, the effect size was found to be in the same order of magnitude (OR 5.14 per 100 g extra cheese per day). Our results on meat and cheese are in line with studies showing that Mediterranean Diet, that comprises few animal products, reduces GD incidence [37, 38]. Furthermore, they are supported by previous work.

Table 4 Associations between selected indicators for animal product intake and gestational diabetes status as estimated by unadjusted and adjusted logistic regression modelling (odds ratios and 95% confidence intervals; n = 2800)

| Characteristic                  | OR (95%CI) unadjusted | p value | OR (95%CI) adjusted* | p value |
|--------------------------------|-----------------------|---------|----------------------|---------|
| Ratio haem to non-haem iron (maternal diet; × 1/10) | 1.80 (1.00, 3.23)     | 0.05    | 1.61 (0.84, 3.07)    | 0.15    |
| Ratio animal to plant protein (maternal diet) | 1.77 (1.14, 2.73)     | 0.01    | 1.78 (1.12, 2.82)    | 0.02    |
| Arachidonic acid (maternal plasma; 1% wt)b | 1.72 (1.18, 2.51)     | 0.01    | 1.72 (1.17, 2.52)    | 0.01    |

Participants with diabetes before pregnancy were excluded from the logistic regression models
Statistically significant p-values are written in bold
*a Covariates included in the model: mother’s age at delivery and gravidity
*b Data available for 1304 pregnant women without GD and 16 with GD
revealing an association between Western dietary patterns and GD [9]. The nature of such associations reserves further investigations. Seeing that vegetarian diets affect gut microbiota [39], a modulation of gut microbiota by dietary patterns could be an appealing possibility.

The relevant time window to adjust dietary pattern and thereby minimise the risk of GD development cannot be extrapolated from the present analysis. Dietary patterns in the past, i.e. long before the pregnancy, might influence health status at the beginning of pregnancy, e.g. by long-term influencing pre-pregnancy BMI that per se is a risk factor for GD [40]. It is, however, also conceivable that dietary pattern changes during pregnancy might still affect the development of GD. In fact, tailored dietary planning—mostly by reducing sugar-sweetened beverages—is an important part of GD management and has the immediate goal of improving glycaemic control [41]. Interestingly, reducing hypercaloric food product intake—including meat products and cheese—may also contribute to energy restriction by obese pregnant women and prevent excessive weight gain during pregnancy together with its long-term consequences ([40, 42]). This is in line with very recent work suggesting that vegetarian diets during pregnancy are associated with reduced gestational weight gain [43].

GD is known to be associated with marked perinatal complications and with adiposity, adverse cardiometabolic outcomes and reduced cognitive ability of the offspring. Our work suggests that pre-conceptional and early pregnancy dietary advice should include recommendations to partially replace meat, meat products and cheese by plant-based products at the earliest opportunity in an attempt to minimise the risk of developing GD.

Supplementary Information  The online version contains supplementary material available at https://doi.org/10.1007/s00394-021-02601-4.

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Author contributions  APS-W and CT designed this analysis of KOALA-data CM-P and MCJMvD processed the original food frequency questionnaire data and controlled their quality. APSW made the present data analysis and wrote the manuscript. CT is the principal investigator of the KOALA-study. All authors were closely involved in revising the article and consented for the final version to be submitted.

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Data availability  All data generated or analysed during the current study are available from the corresponding author on reasonable request.

Declarations

Conflict of interest  None declared.

Ethics approval  The KOALA Birth Cohort Study was approved by the Medical Ethical Committee of Maastricht University/Academic Hospital Maastricht (MUMC+) and was therefore performed in accordance with the ethical standards laid down in the 1964 Declaration of Helsinki and its later amendments.

Consent to participate  All pregnant women (and children’s fathers) signed for written informed consent.

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