Fluid Models of Many-server Queues with Abandonment

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Abstract

We study many-server queues with abandonment in which customers have general service and patience time distributions. The dynamics of the system are modeled using measure-valued processes, to keep track of the residual service and patience times of each customer. Deterministic fluid models are established to provide first-order approximation for this model. The fluid model solution, which is proved to uniquely exists, serves as the fluid limit of the many-server queue, as the number of servers becomes large. Based on the fluid model solution, first-order approximations for various performance quantities are proposed.

Key words and phrases: many-server queue, abandonment, measure valued process, quality driven, efficiency driven, quality and efficiency driven.

1 Introduction

Recently, there has been a great interest in queues with a large number of servers, motivated by applications to telephone call centers. Since a customer can easily hang up after waiting for too long, abandonment is a non-negligible aspect in the study of many-server queues. In our study, a customer can leave the system (without getting service) once has been waiting in queue for more than his patience time. Both patience and service times are modeled using random variables. A recent statistical study by Brown et al. [2] suggests that the exponential assumption on service time distribution, in many cases, is not valid. In fact, the distribution of service times at call centers may be log-normal in some cases as shown in [2]. This emphasizes the need to look at the many-server model with generally distributed service and patience times.
In this paper, we study many-server queues with general patience and service times. The queueing model is denoted by $G/GI/n+GI$. The $G$ represents a general stationary arrival process. The first $GI$ indicates that service times come from a sequences of independent and identically distributed (IID) random variables with a general distribution. The $n$ denotes the number of homogeneous servers. There is an unlimited waiting space, called buffer, where customers wait and can choose to abandon if their patience times expires before their service starts. Again, the patience times of each customer are IID and with a general distribution (the $GI$ after the ‘$+$’ sign).

Useful insights can be obtained by considering a many-server queue in limit regimes where the number $n$ of servers increases along with the arrival rate $\lambda^n$ such that the traffic intensity

$$\rho^n = \frac{\lambda^n}{n\mu} \to \rho$$

as $n \to \infty$, where $\mu$ is the service rate of a single server (in other words, the reciprocal of the mean service time), and $\rho \in [0, \infty)$. Since the abandonment ensures stability, the limit $\rho$ in the above need not to be less than 1. In fact, according to $\rho$, the limit regimes can be divided into three classes, i.e. Efficiency-Driven (ED) regime when $\rho > 1$, Quality-and-Efficiency-Driven (QED) regime when $\rho = 1$ and Quality-Driven (QD) regime when $\rho < 1$. The QED regime is also called Halfin-Whitt regime due to the seminal work Halfin and Whitt [11]. With this motivation, we establish the fluid (also called law of large number) limit for the $G/GI/n+GI$ queue in all the ED, QED and QD limit regimes.

We show that the fluid model has an equilibrium, which yields approximations for various performance quantities. These fluid approximations work pretty well in the ED and QD regime where $\rho$ is not that close to 1, as demonstrated in the numerical experiments of Whitt [28]. However, when $\rho$ is very close (say within 5%) to 1, the fluid approximations lose their accuracy and we shall look at a more refined limit, the diffusion limit, in this case. Diffusion limit is not within the scope of the current paper.

One of the challenges in studying many-server queue with general service (as well as patience) time is that Markovian analysis can not be used. In a system where multiple customers are processed at the same time, such as the many-server queue, how to describe the system becomes an important issue. The number of customers in the system does not give much information since they may all have large remaining service times or all have small remaining service times, and this information can affect future evolution of the system. We choose finite Borel measures on $(0, \infty)$ to describe the system. At any time $t \geq 0$, instead of recording the total number of customers in service (i.e. the number of busy servers), we record all the remaining patience times using measure $Z(t)$. For any Borel set $C \in (0, \infty)$, $Z(t)(C)$ indicates the number of customers in server with remaining service time belongs to $C$ at that time. Similar idea applies for the remaining patience times. We first introduce the virtual buffer, which holds all the customers who have arrived but not yet
scheduled to receive service (assuming they are infinitely patient). We record all the remaining patience times for those in the virtual buffer using finite Borel measure $\mathcal{R}(t)$ on $\mathbb{R} = (-\infty, \infty)$. At time $t \geq 0$, $\mathcal{R}(t)(C)$ indicates the number of customers in the virtual buffer with remaining patience time belongs to the Borel set $C$. The descriptor $(\mathcal{R}(\cdot), \mathbb{Z}(\cdot))$ contains very rich information, almost all information about the system can be recovered from it. Note that a customer with negative remaining patience time has already abandoned. So the actual number of customers in the buffer is

$$Q(t) = \mathcal{R}(t)((0, \infty)) \text{ for all } t \geq 0.$$

More details will be discussed when we rigorously introduce the mathematical model in Section 2.

In literature, another descriptor that keeps track of the ages of customers in service and the ages of customers in waiting have been used, e.g. [15, 28]; The age processes have the advantage of being observable, without requiring future information, though their analysis is often more complicated. Both age and residual descriptions of the system often results in the same steady state insights. In this paper, we focus on residual processes only.

The framework of using measure-valued process has been successfully applied to study models where multiple customers are processed at the same time. Existing works include Gromoll and Kruk [8], Gromoll, Puha and Williams [9] and Gromoll, Robert and Zwart [10], to name a few. Most of these works are on the processor sharing queue and related models where there is no waiting buffer. Recently, Zhang, Dai and Zwart [31, 30] apply the measure-valued process to study the limited processor sharing queue, where only limited number of customers can be served at any time with extra customers waiting in a buffer. Many techniques in this paper closely follows from those developed in [31]. There has been a huge literature on many-server queue and related models since the seminal work by Halfin and Whitt [11]. But there are not many successes with the case where the service time distribution is allowed to be non-exponential. One exception is the work of Reed [25], in which fluid and diffusion limits of the customer-count process of many server queues (without abandonment) are established where few assumptions beyond a first moment are placed on the service time distribution. Later, Puhalskii and Reed [23] extend the aforementioned results to allow noncritical loading, generally distributed service times, and general initial conditions. Jelenković et al. [13] study the many-server queue with deterministic service times; Garmarnik and Momčilović [6] study the model with lattice-valued service times; Puhalskii and Reiman [24] study the model with phase-type service time distributions. Mandelbaum and Momčilović [18] study the virtual waiting time processes, and Kaspi and Ramanan [16] study the fluid limit of measure-valued processes for many-server queues with general service times. For the many-server queue with abandonment, a version of the fluid model have been established as a conjecture in Whitt [27], where a lot of insight was demonstrated, which help greatly in our work. Recently, Kang and Ramanan also worked on the same topic and summarized their result in the technical report [15]. Although
we focus on the same topic, our work uses different methodology from that in [15] and requires less assumptions on the service time distribution. In our work, the only assumption on the service time distribution is continuity, while the service time distribution in [15] is required to have a density and the hazard rate function must be either bounded or lower lower-semicontinuous. From the modeling aspect, our approach mainly based on tracking the “residual” processes, while [15] tracks the “age” processes for studying the queueing model. Also, we propose a quite simple fluid model, which facilitates the analysis. The existence of solution to the fluid model in [15] is proved by showing each fluid limit satisfies the fluid model equations. The current paper proves the existence directly from the definition of the fluid model without invoking fluid limits. In addition, we verify in the end of this paper (c.f. Section 6) that our fluid model is consistent with the special case where both service and patience times are exponentially distributed, as established in Whitt [27] for the ED regime, Garnet et al. [7] for QED regime and Pang and Whitt [21] and Puhalskii [22] for all three regimes. Additional works on many-server queues with abandonment includes Dai, He and Tezcan [4] for phase-type service time distributions and exponential patience time distribution; Zeltyn and Mandelbaum [29] for exponential service time distribution and general patience time distributions; Mandelbaum and Momčilović [19] for both general service time distribution and general patience time distribution. The difference between our work and [19] is that we study the fluid limit of measure-valued processes in all three regimes, and [19] studies the diffusion limit of customer count processes and virtual waiting processes in the QED regime. By assuming a convenient initial condition, [19] does not require a detailed fluid model analysis.

The paper is organized as follows: We begin in Section 2 by formulating the mathematical model of the $G/GI/n+GI$ queue. The dynamics of the system are clearly described by modeling with measure-valued processes; see (2.4) and (2.5). The main results, including a characterization of the fluid model and the convergence of the stochastic processes underlying the $G/GI/n+GI$ queue to the fluid model solution are stated in Section 3. In Section 4, we explore the fluid model and give proofs of all the results on the fluid model. Section 5 is devoted to establishing the convergence of stochastic processes, which includes the proof of pre-compactness and the characterization of the limit as the fluid model solution.

1.1 Notation

The following notation will be used throughout. Let $\mathbb{N}$, $\mathbb{Z}$ and $\mathbb{R}$ denote the set of natural numbers, integers and real numbers respectively. Let $\mathbb{R}_+ = [0, \infty)$. For $a, b \in \mathbb{R}$, write $a^+$ for the positive part of $a$, $\lfloor a \rfloor$ for the integer part, $\lceil a \rceil$ for $\lfloor a \rfloor + 1$, $a \lor b$ for the maximum, and $a \land b$ for the minimum. For any $A \subset \mathbb{R}$, denote $\mathcal{B}(A)$ the collection of all Borel subsets which are subsets of $A$.

Let $\mathcal{M}$ denote the set of all non-negative finite Borel measures on $\mathbb{R}$, and $\mathcal{M}_+$ denote the set of all non-negative finite Borel measures on $(0, \infty)$. To simplify the notation, let us take the convention
that for any Borel set $A \subset \mathbb{R}$, $\nu(A \cap (-\infty, 0]) = 0$ for any $\nu \in \mathcal{M}_+$. Also, by this convention, $\mathcal{M}_+$ is embedded as a subspace of $\mathcal{M}$. For $\nu_1, \nu_2 \in \mathcal{M}$, the Prohorov metric is defined to be

$$d[\nu_1, \nu_2] = \inf \left\{ \epsilon > 0 : \nu_1(A) \leq \nu_2(A^\epsilon) + \epsilon \right\},$$

where $A^\epsilon = \{ b \in \mathbb{R} : \inf_{a \in A} |a - b| < \epsilon \}$. This is the metric that induces the topology of weak convergence of finite Borel measures. (See Section 6 in [1].) For any Borel measurable function $g : \mathbb{R} \to \mathbb{R}$, the integration of this function with respect to the measure $\nu \in \mathcal{M}$ is denoted by $\langle g, \nu \rangle$.

Let $\mathcal{M}_+ \times \mathcal{M}$ denote the Cartesian product. There are a number of ways to define the metric on the product space. For convenience we define the metric to be the maximum of the Prohorov metric between each component. With a little abuse of notation, we still use $d$ to denote this metric.

Let $(E, \pi)$ be a general metric space. We consider the space $D$ of all right-continuous $E$-valued functions with finite left limits defined either on a finite interval $[0, T]$ or the infinite interval $[0, \infty)$. We refer to the space as $D([0, T], E)$ or $D([0, \infty), E)$ depending upon the function domain. The space $D$ is also known as the space of càdlàg functions. For $g(\cdot), g'(\cdot) \in D([0, T], E)$, the uniform metric is defined as

$$\nu_T[g, g'] = \sup_{0 \leq t \leq T} \pi[g(t), g'(t)].$$

However, a more useful metric we will use is the following Skorohod $J_1$ metric,

$$g_T[g, g'] = \inf_{f \in \Lambda_T} (\|f\|_T^2 \lor \nu_T[g, g' \circ f]),$$

where $g \circ f(t) = g(f(t))$ for $t \geq 0$ and $\Lambda_T$ is the set of strictly increasing and continuous mapping of $[0, T]$ onto itself and

$$\|f\|_T^2 = \sup_{0 \leq s < t \leq T} \left| \log \frac{f(t) - f(s)}{t - s} \right|.$$

If $g(\cdot)$ and $g'(\cdot)$ are in the space $D([0, \infty), E)$, the Skorohod $J_1$ metric is defined as

$$g[g, g'] = \int_0^\infty e^{-T(g[g, g'] \wedge 1)}dT.$$  

By saying convergence in the space $D$, we mean the convergence under the Skorohod $J_1$ topology, which is the topology induced by the Skorohod $J_1$ metric [5].

We use “→” to denote the convergence in the metric space $(E, \pi)$, and use “⇒” to denote the convergence in distribution of random variables taking value in the metric space $(E, \pi)$.

## 2 Stochastic Model

In this section, we first describe the $G/GI/n+GI$ queueing system and then introduce a pair of measure-valued processes that capture the dynamics of the system.
There are \( n \) identical servers in the system. Customers arrive according to a general stationary
arrival process (the initial G) with arrival rate \( \lambda \). Let \( a_i \) denote the arrival time of the \( i \)th arriving
customer, \( i = 1, 2, \ldots \). An arriving customer enters service immediately upon arrival if there is a
server available. If all \( n \) servers are busy, the arriving customer waits in a buffer, which has infinite
capacity. Customers are served in the order of their arrival by the first available server. Waiting
customers may also elect to abandon. We assume that each customer has a random patience time.
A customer will abandon immediately when his waiting time in the buffer exceeds his patience time.
Once a customer starts his service, the customer remains until the service is completed. There are
no retrials; abandoning customers leave without affecting future arrivals.

The two GIs in the notation mean that the service times and patience times come from two
independent sequences of iid random variables; these two sequences are assumed to be independent
of the arrival process. Let \( u_i \) and \( v_i \) denote the patience and service time of the \( i \)th arriving
customer, \( i = 1, 2, \ldots \). In many applications such as telephone call centers, customers cannot
see the queue (the case of invisible queues, c.f. [20]), thus do not know the experience of other
customers. In such a case, it is natural to assume that patience times are iid. Denote \( F(\cdot) \) and
\( G(\cdot) \) the distributions for the patience and service times, respectively.

To describe the system using measure-valued process, we first introduce the notion of virtual
buffer. The virtual buffer holds all customers in the real buffer and some of the abandoned cus-
tomers. An abandoned customer continues to wait in the virtual buffer when he first abandons until
it were his turn for service had he not abandoned. At this time, he leaves the virtual buffer. At
any time \( t \geq 0 \), \( R(t) \) denotes a measure in \( \mathcal{M} \) such that \( R(t)(C) \) is the number of customers in the
virtual buffer with remaining patience time in \( C \in \mathcal{B}(\mathbb{R}) \). Please note that this way of modeling
requires \( R(\cdot) \) to be a measure on \( \mathbb{R} \), not just \( (0, \infty) \). It is clear that

\[
Q(t) = R(t)((0, \infty)) \quad \text{and} \quad R(t) = R(t)(\mathbb{R})
\]

(2.1)

represent the number of customers waiting in the real buffer and number of customers in the virtual
buffer, respectively.

We also use a measure to describe the server. At any time \( t \geq 0 \), \( Z(t) \) denotes a measure in \( \mathcal{M}_+ \) such that \( Z(t)(C) \) is the number of customers in service with remaining service time in
\( C \in \mathcal{B}((0, \infty)) \). Different from the virtual buffer, the servers only hold customers with positive
remaining service times, so we only care about the subsets in \( (0, \infty) \). The quantity

\[
Z(t) = Z(t)((0, \infty)),
\]

(2.2)

represents the number of customers in service at any time \( t \geq 0 \).

The measure-valued (taking value in \( \mathcal{M} \times \mathcal{M}_+ \)) stochastic process \((R(\cdot), Z(\cdot))\) serves as the
descriptor for the \( G/GI/n+GI \) queueing model. Before we use it to describe the dynamics of the
system, let us first talk about the initial condition, since the system is allowed to be non-empty initially. The initial state specifies \( R(0) \), the number of customers in the virtual buffer as well as their remaining patience times \( u_i \) and service times \( v_i \), \( i = 1 - R(0), 2 - R(0), \ldots, 0 \). The initial state also specifies \( Z(0) \), the number of customers in service as well as their remaining service times \( v_i \), \( i = 1 - R(0) - Z(0), \ldots, -R(0) \). Briefly, the initial customers are given negative index, in order not to conflict with the index of arriving customers. Those initial customers in the buffer are also assumed to have i.i.d. service times with distribution \( G(\cdot) \). For each \( t \geq 0 \), denote \( E(t) \) the number of customers that has arrived during time interval \((0, t]\). Arriving customers are indexed by \( 1, 2, \ldots \) according to the order of their arrival. By this way of indexing customers, it is clear that the index of the first customer in the virtual buffer at time \( t \geq 0 \) is \( B(t) + 1 \), where

\[
B(t) = E(t) - R(t). \tag{2.3}
\]

Denote \( w_i \) the waiting time of the \( i \)th customers; then \( \tau_i = a_i + w_i \) is the time when the \( i \)th job starts service for all \( i \geq 1 - R(0) \). For \( i < 0 \), \( a_i \) may be a negative number indicating how long the \( i \)th customer had been there by time 0. We will impose some conditions on \( a_i \)'s with \( i < 0 \) later on. Let \( \delta_x \) and \( \delta_{(x,y)} \) denote the Dirac point measure at \( x \in \mathbb{R} \) and \( (x,y) \in \mathbb{R}^2 \), respectively. Denote \( C + x = \{ c + x : x \in C \} \) for any subset \( C \subseteq \mathbb{R} \) and \( C_x = (x, \infty) \). For any subsets \( C, C' \subseteq \mathbb{R} \), let \( C \times C' \) denote the Cartesian product. Using the Dirac measure and the above introduced notations, the evolution of the system can be captured by the following \textit{stochastic dynamic equations}:

\[
\mathcal{R}(t)(C) = \sum_{i=1+B(t)}^{E(t)} \delta_{u_i}(C + t - a_i), \quad \text{for all } C \in \mathcal{B}(\mathbb{R}), \tag{2.4}
\]

\[
\mathcal{Z}(t)(C) = \sum_{i=1-R(0)-Z(0)}^{B(t)} \delta_{v_i}(C + t) + \sum_{i=1-R(0)}^{B(t)} \delta_{(u_i,v_i)}(C_0 + \tau_i - a_i) \times (C + t - \tau_i), \quad \text{for all } C \in \mathcal{B}((0, \infty)), \tag{2.5}
\]

for all \( t \geq 0 \). Denote the total number of customers in the system by

\[
X(t) = Q(t) + Z(t) \quad \text{for all } t \geq 0.
\]

The following \textit{policy constraints} must be satisfied at any time \( t \geq 0 \),

\[
Q(t) = (X(t) - n)^+, \tag{2.6}
\]

\[
Z(t) = (X(t) \land n), \tag{2.7}
\]

where \( n \), as introduced above, denotes the number of servers in the system.
3 Main Results

The main results of this paper contain two parts. The first part is a characterization of the fluid model, including the existence and uniqueness of the fluid model solution, and the equilibrium of the fluid model; these results are summarized in Section 3.1. The second part is the convergence of the stochastic processes to the fluid model solution; this result is stated in Section 3.2.

3.1 Fluid Model

To study the stochastic model, we introduce a deterministic fluid model. To simplify notations, let $F^c(\cdot)$ denote the complement of the patience time distribution $F(\cdot)$, i.e. $F^c(x) = 1 - F(x)$ for all $x \in \mathbb{R}$; the complement of the service time distribution, denoted by $G^c(\cdot)$, is defined in the same way. We introduce the following fluid dynamic equations:

$$
\bar{R}(t)(C_x) = \lambda \int_{t-B(t)}^t F^c(x + t - s)ds, \quad t \geq 0, \quad x \in \mathbb{R},
$$

$$
\bar{Z}(t)(C_x) = \bar{Z}(0)(C_x + t) + \int_0^t F^c \left( \frac{R(s)}{\lambda} \right) G^c(x + t - s)d\bar{B}(s), \quad t \geq 0, \quad x \in (0, \infty),
$$

where $C_x = (x, \infty)$ and $\bar{B}(s) = \lambda s - \bar{R}(s)$. Here, all the time dependent quantities are assumed to be right continuous on $[0, \infty)$ and to have left limits in $(0, \infty)$; furthermore, $\bar{B}(\cdot)$ is a non-decreasing function, and the integral $\int_0^t g(s) d\bar{B}(s)$ is interpreted as the Lebesgue-Stieltjes integral on the interval $[0, t]$. The quantities $\bar{R}(\cdot), \bar{Q}(\cdot), \bar{Z}(\cdot)$ and $\bar{X}(\cdot)$ are defined in the same way as their stochastic counterparts in (2.1), (2.2) and (2). The following policy constraints must be satisfied for all $t \geq 0$,

$$
\bar{Q}(t) = (\bar{X}(t) - 1)^+, \quad (3.3)
$$

$$
\bar{Z}(t) = (\bar{X}(t) \wedge 1). \quad (3.4)
$$

The fluid dynamic equations (3.1) and (3.2) and the policy constraints (3.3) and (3.4) define a fluid model, which is denoted by $(\lambda, F, G)$.

Denote $(\bar{R}_0, \bar{Z}_0) = (\bar{R}(0), \bar{Z}(0))$ the initial condition of the fluid model. For the convenience of notations, also denote $\bar{Q}_0 = \bar{Q}(0)$, $\bar{Z}_0 = \bar{Z}(0)$ and $\bar{X}_0 = \bar{Q}_0 + \bar{Z}_0$. We need to require that the initial condition satisfies the dynamic equations and the policy constraints, i.e.

$$
\bar{R}_0(C_x) = \lambda \int_0^{\bar{X}_0} F^c(x + s)ds, \quad x \in \mathbb{R},
$$

$$
\bar{Q}_0 = (\bar{X}_0 - 1)^+, \quad (3.5)
$$

$$
\bar{Z}_0 = (\bar{X}_0 \wedge 1). \quad (3.6)
$$
We also require that
\[ \bar{Z}_0(\{0\}) = 0, \quad (3.8) \]
which means that nobody with remaining service time 0 stays in the server. We call any element
\((\bar{R}_0, \bar{Z}_0) \in M \times M_+\) a valid initial condition if it satisfies (3.5)–(3.8).

We call \((\bar{R}(\cdot), \bar{Z}(\cdot)) \in D([0, \infty), M \times M_+)\) a solution to the fluid model \((\lambda, F, G)\) with a valid initial condition \((\bar{R}_0, \bar{Z}_0)\) if it satisfies the fluid dynamic equations (3.1) and (3.2) and the policy constraints (3.3) and (3.4).

Denote \(\mu\) the reciprocal of first moment of the service time distribution \(G(\cdot)\). Let
\[ MF = \inf\{x \geq 0 : F(x) = 1\}. \quad (3.9) \]
By the right continuity, it is clear that \(F(x) < 1\) for all \(x < MF\) and \(F(x) = 1\) for all \(x \geq MF\).

If the patience time distribution \(F(\cdot)\) has a density \(f(\cdot)\), then define the hazard rate \(h_F(\cdot)\) of the distribution \(F(\cdot)\) by
\[ h_F(x) = \begin{cases} \frac{f(x)}{1-F(x)} & x < MF, \\ 0 & x \geq MF. \end{cases} \]

**Theorem 3.1 (Existence and Uniqueness).** Assume the service time distribution satisfies both that
\[ G(\cdot) \text{ is continuous,} \quad (3.10) \]
and that
\[ 0 < \mu < \infty. \quad (3.11) \]
Assume the patience time distribution satisfies either that
\[ F(\cdot) \text{ is Lipschitz continuous,} \quad (3.12) \]
or that \(F(\cdot)\) has a density \(f(\cdot)\) such that the hazard rate is bounded, i.e.
\[ \sup_{x \in [0, \infty)} h_F(x) < \infty. \quad (3.13) \]
There exists a unique solution to the fluid model \((\lambda, F, G)\) for any valid initial condition \((\bar{R}_0, \bar{Z}_0)\).

The above theorem provides the foundation to further study the fluid model. A key property is that the fluid model has an equilibrium state. An equilibrium state is defined as the following:

**Definition 3.1.** An element \((\bar{R}_\infty, \bar{Z}_\infty) \in M \times M_+\) is called an equilibrium state for the fluid model \((\lambda, F, G)\) if the solution to the fluid model with initial condition \((\bar{R}_\infty, \bar{Z}_\infty)\) satisfies
\[ (\bar{R}(t), \bar{Z}(t)) = (\bar{R}_\infty, \bar{Z}_\infty) \text{ for all } t \geq 0. \]
This definition says that if a fluid model solution starts from an equilibrium state, it will never change in the future. To present the result about equilibrium state, we need to introduce some more notation. For the service time distribution function $G(\cdot)$ on $\mathbb{R}_+$, the associated *equilibrium* distribution is given by

$$G_e(x) = \mu \int_0^x G^c(y)dy, \quad \text{for all } x \geq 0.$$  

**Theorem 3.2.** Assume the conditions in Theorem 3.1. The state $(\bar{R}_\infty, \bar{Z}_\infty)$ is an equilibrium state of the fluid model $(\lambda, F, G)$ if and only if it satisfies

$$\bar{R}_\infty(C_x) = \lambda \int_0^w F^c(x + s)ds, \quad x \in \mathbb{R}, \quad (3.14)$$

$$\bar{Z}_\infty(C_x) = \min(\rho, 1)[1 - G_e(x)], \quad x \in (0, \infty), \quad (3.15)$$

where $w$ is a solution to the equation

$$F(w) = \max\left(\frac{\rho - 1}{\rho}, 0\right). \quad (3.16)$$

**Remark 3.1.** If equation (3.16) has multiple solutions, then the equilibrium is not unique (any solution $w$ gives an equilibrium). If the equation has a unique solution (for example when $F(\cdot)$ is strictly increasing), then the equilibrium state is unique.

The quantity $w$ is interpreted to be the offered waiting time for an arriving customer. If his patience time exceeds $w$, he will not abandon. Thus, the probability of his abandonment is given by $F(w)$, which is equal to $(\rho - 1)/\rho$ when $\rho > 1$; the latter quantity is the fraction of traffic that has to be discarded due to the overloading. From (3.14), $\bar{R}_\infty(C_x) = \lambda w$ for $x \leq -w$. Thus, the average number of customers in the virtual buffer is

$$\bar{R}_\infty = \bar{R}_\infty(\mathbb{R}) = \lambda w,$$

which is consistent with the Little’s law. From (3.15), the average number of busy servers is

$$\bar{Z}_\infty = \bar{Z}_\infty((0, \infty)) = \min(\rho, 1),$$

which is intuitively clear. These observations and interpretations were first made by Whitt [28], where approximation formulas based on a conjectured fluid model were also given, and were compared with extensive simulation results. The approximation formulas derived from our fluid model is consistent with those formulas in Whitt [28].

### 3.2 Convergence of Stochastic Models

We consider a sequence of queueing systems indexed by the number of servers $n$, with $n \to \infty$. Each model is defined in the same way as in Section 2. The arrival rate of each model is assumed
be to proportional to $n$. To distinguish models with different indices, quantities of the $n$th model are accompanied with superscript $n$. Each model may be defined on a different probability space $(\Omega^n, \mathcal{F}^n, \mathbb{P}^n)$. Our results concern the asymptotic behavior of the descriptors under the fluid scaling, which is defined by

$$
\bar{R}^n(t) = \frac{1}{n} R^n(t), \quad \bar{Z}^n(t) = \frac{1}{n} Z^n(t),
$$

(3.17)

for all $t \geq 0$. The fluid scaling for the arrival process $E^n(\cdot)$ is defined in the same way, i.e.

$$
\bar{E}^n(t) = \frac{1}{n} E^n(t),
$$

for all $t \geq 0$. We assume that

$$
\bar{E}^n(\cdot) \Rightarrow \lambda \cdot \text{ as } n \to \infty.
$$

(3.18)

Since the limit is deterministic, the convergence in distribution in (3.18) is equivalent to convergence in probability; namely, for each $T > 0$ and each $\epsilon > 0$,

$$
\lim_{n \to \infty} \mathbb{P}^n \left( \sup_{0 \leq t \leq T} |\bar{E}^n(t) - \lambda t| > \epsilon \right) = 0.
$$

Denote $\nu^n_F$ and $\nu^n_G$ the probability measures corresponding to the patience time distribution $F^n$ and the service time distribution $G^n$, respectively. Assume that as $n \to \infty$,

$$
\nu_F^n \to \nu_F, \quad \nu_G^n \to \nu_G,
$$

(3.19)

where $\nu_F$ and $\nu_G$ are some probability measures with associated distribution functions $F$ and $G$. Also, the following initial condition will be assumed:

$$
(\bar{R}^n(0), \bar{Z}^n(0)) \Rightarrow (\bar{R}_0, \bar{Z}_0) \quad \text{as } n \to \infty,
$$

(3.20)

where, almost surely, $(\bar{R}_0, \bar{Z}_0)$ is a valid initial condition and

$$
\bar{R}_0 \text{ and } \bar{Z}_0 \text{ has no atoms.}
$$

(3.21)

**Theorem 3.3.** In addition to the assumptions (3.10)–(3.13) in Theorem 3.1, if the sequence of many-server queues satisfies (3.18)–(3.21), then

$$
(\bar{R}^n(\cdot), \bar{Z}^n(\cdot)) \Rightarrow (\bar{R}(\cdot), \bar{Z}(\cdot)) \quad \text{as } n \to \infty,
$$

where, almost surely, $(\bar{R}(\cdot), \bar{Z}(\cdot))$ is the unique solution to the fluid model $(\lambda, F, G)$ with initial condition $(\bar{R}_0, \bar{Z}_0)$.

## 4 Properties of the Fluid Model

In this section, we analyze the proposed fluid model and establish some basic properties of the fluid model solution. The proof of Theorem 3.1 for existence and uniqueness and the proof of Theorem 3.2 for characterization of the equilibrium will be presented in Section 4.1 and Section 4.2, respectively.
4.1 Existence and Uniqueness of Fluid Model Solutions

We first present some calculus on the fluid dynamic equations (3.1) and (3.2), which define the fluid model. It follows from (3.1) that

\[ \bar{Q}(t) = \bar{R}(t)(C_0) = \lambda \int_{t - \frac{R(t)}{\lambda}}^{t} F^c(t - s)ds = \lambda \int_{0}^{\frac{R(t)}{\lambda}} F^c(s)ds. \]

Let

\[ F_d(x) = \int_{0}^{x} [1 - F(y)]dy \quad \text{for all } x \geq 0. \]

Please note that the density of \( F_d(\cdot) \) is not scaled by the mean of \( F(\cdot) \). Thus, this is not exactly the equilibrium distribution associated with \( F(\cdot) \). In fact, we do not need the mean \( N_F = \int_{0}^{\infty} [1 - F(y)]dy \) (4.1) to be finite. Now we have

\[ \frac{\bar{Q}(t)}{\lambda} = F_d \left( \frac{\bar{R}(t)}{\lambda} \right). \] (4.2)

It follows from (3.2) that

\[ \bar{Z}(t) = \bar{Z}(t)(C_0) = \bar{Z}_0(C_0 + t) + \lambda \int_{0}^{t} F^c \left( \frac{\bar{R}(s)}{\lambda} \right) G^c(t - s)ds \]

\[ - \int_{0}^{t} F^c \left( \frac{\bar{R}(s)}{\lambda} \right) G^c(t - s)d\bar{R}(s). \]

Note that by (4.2), \( d\bar{Q}(s) = F^c \left( \frac{\bar{R}(s)}{\lambda} \right) d\bar{R}(s) \). So

\[ \bar{Z}(t) = \bar{Z}_0(C_0 + t) + \frac{\lambda}{\mu} \int_{0}^{t} F^c \left( \frac{\bar{R}(s)}{\lambda} \right) dG^c(t - s) - \int_{0}^{t} G^c(t - s)d\bar{Q}(s). \]

Performing change of variable and integration by parts, we have

\[ \bar{Z}(t) = \bar{Z}_0(C_t) + \frac{\lambda}{\mu} \int_{0}^{t} F^c \left( \frac{\bar{R}(t - s)}{\lambda} \right) dG^c(s) \]

\[ - \bar{Q}(t)G^c(0) + \bar{Q}(0)G^c(t) + \int_{0}^{t} \bar{Q}(t - s)dG(s). \] (4.3)

We wish to represent the term \( F^c \left( \frac{\bar{R}(\cdot)}{\lambda} \right) \) using \( \bar{Q}(\cdot) \). Recall \( M_F \) and \( N_F \), which are defined in (3.9) and (4.1), respectively. It is clear that \( F_d(x) \) is strictly monotone for \( x \in [0, M_F] \). Thus, \( F_d^{-1}(y) \) is well defined for each \( y \in [0, N_F] \). We define \( F_d^{-1}(y) = M_F \) for all \( y \geq N_F \). Thus, (4.2) implies that

\[ F^c \left( \frac{\bar{R}(t)}{\lambda} \right) = F^c \left( \frac{F_d^{-1}(\bar{Q}(t))}{\lambda} \right). \] (4.4)
Note that $G^c(0) = 1$ by assumption (3.10). Combining (3.3), (3.4), (4.3), and (4.4), we obtain

$$\begin{align*}
\tilde{X}(t) &= \tilde{Z}_0(C_t) + \tilde{Q}_0G^c(t) \\
&\quad + \frac{\lambda}{\mu} \int_0^t F^c\left( F_d^{-1}(\frac{\tilde{X}(t-s) - 1}{\lambda}) \right) dG_e(s) \\
&\quad + \int_0^t (\tilde{X}(t-s) - 1)^+ dG(s).
\end{align*}$$

Now, introduce

$$H(x) = \begin{cases} F^c(F_d^{-1}(\frac{x}{\lambda})) & \text{if } 0 \leq x < \lambda, \\
0 & \text{if } x \geq \lambda,
\end{cases}$$

and $\zeta(\cdot) = \tilde{Z}_0(C_0 + \cdot) + \tilde{Q}_0G^c(\cdot)$. It then follows that

$$\begin{align*}
\tilde{X}(t) &= \zeta_0(t) + \rho \int_0^t H((\tilde{X}(t-s) - 1)^+)dG_e(s) + \int_0^t (\tilde{X}(t-s) - 1)^+dG(s).
\end{align*}$$

Please note that $\zeta_0(\cdot)$ depends only on the initial condition and $H(\cdot)$ is a function defined by the arrival rate $\lambda$ and the patience time distribution $F(\cdot)$. The equation (4.5) serves as a key to the analysis of the fluid model.

**Proof of Theorem 3.1.** We first prove the existence. Given a valid initial condition $(\tilde{R}_0, \tilde{Z}_0)$ (i.e. an element in $M \times M_+$ that satisfies (3.5)–(3.8)), we now construct a solution $(\tilde{R}(\cdot), \tilde{Z}(\cdot))$ to the fluid model $(\lambda, F, G)$ with this initial condition. If the patience time distribution $F(\cdot)$ is Lipschitz continuous, then it is clear that $H(\cdot)$ is also Lipschitz continuous; if $F(\cdot)$ has a density, then the function $H(\cdot)$ is differentiable and has derivative

$$H'(x) = -f(y) \frac{1}{F_d'(y)} = -\frac{f(y)}{1 - F(y)} = -h_F(y),$$

on the interval $(0, \lambda N_F)$ if $y = F_d^{-1}(x)$ and $H(x) = 0$ for all $x \geq \lambda N_F$. By condition (3.13),

$$\sup_{0 < x < \lambda N_F} |H'(x)| = \sup_{y \in [0, M_F]} h_F(y),$$

which implies that $H(\cdot)$ is Lipschitz continuous. It follows from Lemma A.1 that the equation (4.5) has a unique solution $\tilde{X}(\cdot)$. Denote $\tilde{Q}(t) = (\tilde{X}(t) - 1)^+$. We first claim that $\tilde{Q}(t)/\lambda \leq N_F$ for all $t \geq 0$. The claim is automatically true if $N_F = \infty$. Now, let us consider the case where $N_F < \infty$. Since $(\tilde{R}_0, \tilde{Z}_0)$ is a valid initial condition, $\tilde{Q}(0)/\lambda \leq N_F$. Suppose there exists $t_1 > 0$ such that $\tilde{Q}(t_1)/\lambda > N_F$. Let $t_0 = \sup\{s : \tilde{Q}(s)/\lambda \leq N_F, s \leq t_1\}$. So we have that lim$_{t \to t_0} \tilde{Q}(t)/\lambda \leq N_F$, since $\tilde{Q}(\cdot)$ has left limit. Let $\delta = (\tilde{Q}(t_1)/\lambda - N_F)/4$ and pick $t_\delta \in [t_0 - \delta, t_0]$ such that $\tilde{Q}(t_\delta)/\lambda \leq N_F + \delta$. By Lemma A.2,

$$\frac{\tilde{Q}(t')}{\lambda} - \frac{\tilde{Q}(t)}{\lambda} \leq \int_t^{t'} F^c(F_d^{-1}(\frac{\tilde{Q}(s)}{\lambda}))ds$$

\[\tag{4.6}\]
for any $t < t'$. This gives that

$$
\frac{Q(t_1)}{\lambda} \leq \frac{Q(t_0)}{\lambda} + \int_{t_0}^{t_1} \left[1 - F(F_d^{-1}\left(\frac{Q(s)}{\lambda}\right))\right]ds
$$

$$
\leq N_F + \delta + \int_{t_0}^{t_1} 1ds + \int_{t_0}^{t_1} 0ds
$$

$$
\leq N_F + 2\delta < \frac{Q(t_1)}{\lambda},
$$

which is a contradiction. This proves the claim. Let

$$
\tilde{Z}(t) = \min(\tilde{X}(t), 1),
$$

$$
\tilde{R}(t) = \lambda F^{-1}_d\left(\frac{Q(t)}{\lambda}\right),
$$

$$
\tilde{B}(t) = \lambda t - \tilde{R}(t),
$$

for all $t \geq 0$. Next, we claim that the process $\tilde{B}(\cdot)$ is non-decreasing. To prove this claim, it is enough show that

$$
F_d^{-1}\left(\frac{Q(t')}{\lambda}\right) - F_d^{-1}\left(\frac{Q(t)}{\lambda}\right) \leq t' - t \tag{4.7}
$$

for any $t \leq t'$. Since $F_d^{-1}$ is a non-decreasing function, the inequality holds trivially when $\tilde{Q}(t') \leq \tilde{Q}(t)$. We now focus on the case where $\tilde{Q}(t') > \tilde{Q}(t)$. Note that the function $F_d^{-1}(\cdot)$ is convex, since the derivative is non-decreasing. This together with (4.6) implies that

$$
F_d^{-1}\left(\frac{Q(t')}{\lambda}\right) \leq F_d^{-1}\left(\frac{Q(t)}{\lambda}\right) + F_d^{-1}'\left(\frac{Q(t)}{\lambda}\right) \int_{t}^{t'} F^{-1}(F_d^{-1}\left(\frac{Q(s)}{\lambda}\right))ds.
$$

If $\tilde{Q}(t) \leq \tilde{Q}(s)$ for all $s \in [t, t']$, then due to the fact that $F^{-1}(F_d^{-1}(\cdot))$ is non-increasing, we have

$$
F_d^{-1}\left(\frac{Q(t')}{\lambda}\right) \leq F_d^{-1}\left(\frac{Q(t)}{\lambda}\right) + \frac{1}{F^{-1}(F_d^{-1}(\frac{Q(t)}{\lambda}))} F^{-1}(F_d^{-1}\left(\frac{Q(t)}{\lambda}\right))(t' - t)
$$

$$
\leq F_d^{-1}\left(\frac{Q(t)}{\lambda}\right) + t' - t,
$$

which gives (4.7); otherwise, let $t^* \in (t, t')$ be the point where $\tilde{Q}(\cdot)$ achieves minimum. Since $\tilde{Q}(t) > \tilde{Q}(t^*)$, we have

$$
F_d^{-1}\left(\frac{Q(t^*)}{\lambda}\right) - F_d^{-1}\left(\frac{Q(t)}{\lambda}\right) \leq t^* - t.
$$

Since $\tilde{Q}(t^*) \leq \tilde{Q}(s)$ for all $s \in [t^*, t']$, by the same reasoning in the above, we also have

$$
F_d^{-1}\left(\frac{Q(t')}{\lambda}\right) \leq F_d^{-1}\left(\frac{Q(t^*)}{\lambda}\right) + t' - t^*.
$$

The above two inequalities also leads to (4.7). So the claim is proved. We now construct a fluid model solution by letting

$$
\bar{R}(t)(C_x) = \lambda \int_{t-\frac{\bar{R}(t)}{\lambda}}^{t} F^c(x + t - s)ds,
$$

$$
\bar{Z}(t)(C_x) = \bar{Z}_0(C_x + t) + \int_{0}^{t} F^c\left(\frac{\bar{R}(s)}{\lambda}\right)G^c(x + t - s)d\bar{B}(s),
$$

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for all \( t \geq 0 \). It is clear that the above defined \((\bar{\mathcal{R}}(\cdot), \bar{Z}(\cdot))\) satisfies the fluid dynamic equations (3.1) and (3.2) and constraints (3.3) and (3.4). So we conclude that \((\bar{\mathcal{R}}(\cdot), \bar{Z}(\cdot))\) is a fluid model solution.

It now remains to show the uniqueness. Suppose there is another solution to the fluid model \((\lambda, F, G)\) with initial condition \((\bar{R}_0, \bar{Z}_0)\), denoted by \((\bar{R}^\dagger(\cdot), \bar{Z}^\dagger(\cdot))\). Similarly, denote

\[
\begin{align*}
\bar{R}^\dagger(t) &= \bar{R}^\dagger(\mathbb{R}), \\
\bar{Z}^\dagger(t) &= \bar{Z}^\dagger((0, \infty)),
\end{align*}
\]

for all \( t \geq 0 \). It must satisfy the fluid dynamic equations (3.1) and (3.2) and constraints (3.3) and (3.4). For all \( t \geq 0 \), let

\[
\bar{Q}^\dagger(t) = \lambda F_d(\frac{\bar{R}^\dagger(t)}{\lambda}).
\]

According to the algebra at the beginning of Section 4.1, \( \bar{X}^\dagger(\cdot) \) must also satisfy equation (4.5). By the uniqueness of the solution to the equation (4.5) in Lemma A.1,

\[
\bar{X}^\dagger(t) = \bar{X}(t) \quad \text{for all } t \geq 0.
\]

This implies that \( \bar{R}^\dagger(t) = \bar{R}(t) \). By the dynamic equations (3.1) and (3.2), we must have that

\[
(\bar{R}^\dagger(t), \bar{Z}^\dagger(t)) = (R(t), Z(t)) \quad \text{for all } t \geq 0.
\]

This completes the proof. \( \square \)

### 4.2 Equilibrium State of the Fluid Model Solution

In this section, we first intuitively explain what an equilibrium should be. Then we rigorously prove it in Theorem 3.2. To provide some intuition, note that in the equilibrium, by equation (3.1), one should have

\[
\bar{\mathcal{R}}_\infty(C_x) = \lambda \int_0^{\bar{\mathcal{R}}_\infty/\lambda} F^c(x + s)ds,
\]

for the buffer. This immediately implies that

\[
\bar{\mathcal{R}}_\infty(C_x) = \lambda[F_d(x + \frac{\bar{\mathcal{R}}_\infty}{\lambda}) - F_d(x)].
\]

So the rate at which customers leave the buffer due to abandonment is:

\[
\lim_{x \to 0} \frac{\bar{\mathcal{R}}_\infty(C_0) - \bar{\mathcal{R}}_\infty(C_x)}{x} = \lambda F\left(\frac{\bar{R}_\infty}{\lambda}\right).
\]

In the equilibrium, intuitively, the number of customers in service should not change and the distribution for the remaining service time should be the equilibrium distribution \( G_e(\cdot) \), i.e.

\[
\bar{Z}_\infty(C_x) = \bar{Z}_\infty[1 - G_e(x)].
\]
The rate at which customers depart from the server is:

\[ \lim_{x \to 0} \frac{\bar{Z}_\infty(C_0) - \bar{Z}_\infty(C_x)}{x} = \bar{Z}_\infty \mu. \]

The arrival rate must be equal to the summation of the departure rate from server (due to service completion) and the one from buffer (due to abandonment), i.e.

\[ \lambda = \lambda F\left(\frac{\bar{R}_\infty}{\lambda}\right) + \bar{Z}_\infty \mu. \]  \hfill (4.8)

It follows directly from (4.2) that

\[ \bar{Q}_\infty = \lambda F_d\left(\frac{\bar{R}_\infty}{\lambda}\right). \]  \hfill (4.9)

If \( \bar{R}_\infty > 0 \), then according to (4.9) we have \( \bar{Q}_\infty > 0 \). Thus \( \bar{Z}_\infty = 1 \) according to policy constraints. By (4.8), \( \rho > 1 \) and \( \frac{\bar{R}_\infty}{\lambda} \) is a solution to the equation \( F(w) = \frac{\rho - 1}{\rho} \). If \( \bar{R}_\infty = 0 \), then according to (4.8) we have \( \rho = \bar{Z}_\infty \leq 1 \). In summary, we have that

\[ \bar{Q}_\infty = \lambda F_d(w), \]

\[ \bar{Z}_\infty = \min(\rho, 1), \]

where \( w \) is a solution to the equation \( F(w) = \max\left(\frac{\rho - 1}{\rho}, 0\right) \). This is consistent with the one in [28], which is derived from a conjecture of a fluid model. Now, we rigorously prove this result.

**Proof of Theorem 3.2.** If \((\bar{R}_\infty, \bar{Z}_\infty)\) is an equilibrium state, then according to the definition, it must satisfies

\[ \bar{R}_\infty(C_x) = \lambda \int_{t-\frac{\bar{R}_\infty}{\lambda}}^{t} F^c(x + t - s)ds, \quad t \geq 0, \]  \hfill (4.10)

\[ \bar{Z}_\infty(C_x) = \bar{Z}_\infty(C_x + t) + \int_{0}^{t} F^c\left(\frac{\bar{R}_\infty}{\lambda}\right) G_c(x + t - s)d\lambda s, \quad t \geq 0. \]  \hfill (4.11)

It follows from (4.11) that

\[ \bar{Z}_\infty(C_x) - \bar{Z}_\infty(C_x + t) = \rho F^c\left(\frac{\bar{R}_\infty}{\lambda}\right) \mu \int_{0}^{t} G_c(x + t - s)ds \]

\[ = \rho F^c\left(\frac{\bar{R}_\infty}{\lambda}\right) [G_c(x + t) - G_c(x)], \quad t \geq 0. \]

Taking \( t \to \infty \), one has

\[ \bar{Z}_\infty(C_x) = \rho F^c\left(\frac{\bar{R}_\infty}{\lambda}\right) G_c^e(x). \]  \hfill (4.12)

Thus \( \bar{Z}_\infty = \rho F^c\left(\frac{\bar{R}_\infty}{\lambda}\right) \). According to (4.2), we have that

\[ \bar{Q}_\infty = \lambda F_d\left(\frac{\bar{R}_\infty}{\lambda}\right). \]
First assume that \( \bar{R}_\infty > 0 \). Then \( \bar{Q}_\infty > 0 \), and thus \( \bar{Z}_\infty = 1 \) by the policy constraints (3.3) and (3.4). Therefore, \( \rho F'(\frac{\bar{R}_\infty}{\bar{X}}) = 1 \), which implies that \( F'(\frac{\bar{R}_\infty}{\bar{X}}) = \frac{\rho - 1}{\rho} \) and \( \rho > 1 \). Now assume that \( \bar{R}_\infty = 0 \). Then \( \bar{Z}_\infty = \rho \), which must be less than or equal to 1 by the policy constraints. Summarizing the cases where \( \rho > 1 \) and \( \rho \leq 1 \), we have that the equilibrium state must satisfy (3.14)–(3.16).

If a state \((\bar{R}_\infty, \bar{Z}_\infty)\) satisfies (3.14)–(3.16), then let
\[
(\bar{R}(t), \bar{Z}(t)) = (\bar{R}_\infty, \bar{Z}_\infty),
\]
for all \( t \geq 0 \). If \( \rho \leq 1 \), then \( \bar{R}(\cdot) \equiv 0 \) and \( \bar{Z}(\cdot) \equiv \rho \); if \( \rho > 1 \), then \( \bar{R}(\cdot) \equiv \lambda w \) and \( \bar{Z}(\cdot) \equiv 1 \), where \( w \) is a solution to equation (3.16). It is easy to check that \((\bar{R}(\cdot), \bar{Z}(\cdot))\) is a fluid model solution in both cases. So by definition, the state \((\bar{R}_\infty, \bar{Z}_\infty)\) is an equilibrium state.

5 Fluid Approximation of the Stochastic Models

Similar to (2.3), let
\[
B^n(t) = E^n(t) - R^n(t).
\]
(5.1)

It follows from (2.4) and (2.5) that the dynamics for the fluid scaled processes can be written as
\[
\bar{R}^n(t)(C) = \frac{1}{n} \sum_{i=B^n(t)+1}^{E^n(t)} \delta_{u_1^n}(C + t - a^n_i), \quad \text{for all } C \in \mathcal{B}(\mathbb{R}),
\]
(5.2)
\[
\bar{Z}^n(t)(C) = \bar{Z}^n(s)(C + t - s)
\]
\[
+ \frac{1}{n} \sum_{i=B^n(s)+1}^{B^n(t)} \delta_{(u_i^n, v_i^n)}(C_0 + \tau^n_i - a^n_i) \times (C + t - \tau^n_i), \quad \text{for all } C \in \mathcal{B}((0, \infty)),
\]
(5.3)

for all \( 0 \leq s \leq t \).

5.1 Precompactness

We first establish the following precompactness for the sequence of fluid scaled stochastic processes \(\{ (\bar{R}^n(\cdot), \bar{Z}^n(\cdot)) \} \).

**Theorem 5.1.** Assume (3.18)–(3.21). The sequence of the fluid scaled stochastic processes \(\{ (\bar{R}^n(\cdot), \bar{Z}^n(\cdot)) \}_{n \in \mathbb{N}} \) is precompact as \( n \to \infty \); namely, for each subsequence \(\{ (\bar{R}^{n_k}(\cdot), \bar{Z}^{n_k}(\cdot)) \}_{n_k} \) with \( n_k \to \infty \), there exists a further subsequence \(\{ (\bar{R}^{n_{kj}}(\cdot), \bar{Z}^{n_{kj}}(\cdot)) \}_{n_{kj}} \) such that
\[
(\bar{R}^{n_{kj}}(\cdot), \bar{Z}^{n_{kj}}(\cdot)) \Rightarrow (\bar{R}(\cdot), \bar{Z}(\cdot)) \quad \text{as } j \to \infty,
\]
for some \((\bar{R}(\cdot), \bar{Z}(\cdot)) \in \mathcal{D}([0, \infty), M \times M_+). \)
The remaining of this section is devoted to proving the above theorem. By Theorem 3.7.2 in [5], it suffices to verify \((a)\) the compact containment property; Lemma 5.1 and \((b)\) the oscillation bound, Lemma 5.4 below.

5.1 Compact Containment

A set \(K \subset M\) is relatively compact if \(\sup_{\xi \in K} \xi(\mathbb{R}) < \infty\), and there exists a sequence of nested compact sets \(A_j \subset \mathbb{R}\) such that \(\bigcup A_j = \mathbb{R}\) and

\[
\lim_{j \to \infty} \sup_{\xi \in K} \xi(A_j^c) = 0,
\]

where \(A_j^c\) denotes the complement of \(A_j\); see [14], Theorem A7.5. The first major step to prove Theorem 5.1 is to establish the following compact containment property.

**Lemma 5.1.** Assume (3.18)–(3.21). Fix \(T > 0\). For each \(\eta > 0\) there exists a compact set \(K \subset M\) such that

\[
\liminf_{n \to \infty} P^n \left( (\bar{R}^n(t), \bar{Z}^n(t)) \in K \times K \text{ for all } t \in [0,T] \right) \geq 1 - \eta.
\]

To prove this result, we first need to establish some bound estimations. For the convenience of notation, denote \(\bar{E}^n(s,t) = \bar{E}^n(t) - \bar{E}^n(s)\) for any \(0 \leq s \leq t\). Fix \(T > 0\). It follows immediately from condition (3.18) that for each \(\epsilon > 0\) there exists an \(n_0\) such that when \(n > n_0\),

\[
P^n \left( \sup_{0 \leq s < t \leq T} |\bar{E}^n(s,t) - \lambda(t-s)| < \epsilon \right) \geq 1 - \epsilon.
\]

(5.4)

To facilitate some arguments later on, we derive the following result from the above inequality.

**Lemma 5.2.** Fix \(T > 0\). There exists a function \(\epsilon_E(\cdot)\), with \(\lim_{n \to \infty} \epsilon_E(n) = 0\) such that

\[
P^n \left( \sup_{0 \leq s < t \leq T} |\bar{E}^n(s,t) - \lambda(t-s)| < \epsilon_E(n) \right) \geq 1 - \epsilon_E(n),
\]

for each \(n \geq 0\).

The derivation of the above lemma from (5.4) follows the same as the proof of Lemma 5.1 in [31]. We omit the proof for brevity. Based on the above lemma, we construct the following event,

\[
\Omega^n_E = \{ \sup_{t \in [0,T]} |\bar{E}^n(s,t) - \lambda(t-s)| < \epsilon_E(n) \}.
\]

We have that on this event, the arrival process is regular, i.e. \(\bar{E}^n(s,t)\) is “close” to \(\lambda(t-s)\). And this event has “large” probability, i.e.

\[
\lim_{n \to \infty} P^n \left( \Omega^n_E \right) = 1.
\]

(5.6)
Proof of Lemma 5.1. By the convergence of the initial condition \((3.20)\), for any \(\epsilon > 0\), there exists a relatively compact set \(K_0 \subset M\) such that

\[
\liminf_{n \to \infty} \mathbb{P}^n\left( \bar{R}^n(0) \in K_0 \text{ and } \bar{Z}^n(0) \in K_0 \right) > 1 - \epsilon. \tag{5.7}
\]

Denote the event in the above probability by \(\Omega^n_0\). On this event, by the definition of relatively compact set in the space \(M\), there exists a function \(\kappa_0(\cdot)\) with \(\lim_{x \to \infty} \kappa_0(x) = 0\) such that

\[
\bar{R}^n(0)(C_x) \leq \kappa_0(x), \quad \bar{Z}^n(0)(C_x) \leq \kappa_0(x), \tag{5.8}
\]

and

\[
\bar{R}^n(0)(C_x^-) \leq \kappa_0(x), \tag{5.9}
\]

for all \(x \geq 0\), where \(C_x^- = (-\infty, -x)\) for any \(y \in \mathbb{R}\). (Remember that \(\bar{Z}^n(0)\) is a measure on \((0, \infty)\), so we do not need to consider its measure of \(C_x^-\).) It is clear that on the event \(\Omega^n_E \cap \Omega^n_0\), for any \(t \leq T\) and all large \(n\),

\[
\bar{R}^n(t)(\mathbb{R}) \leq \sup_n \bar{R}^n(0)(\mathbb{R}) + 2\lambda T, \quad \bar{Z}^n(t)((0, \infty)) \leq 1,
\]

where the last inequality is due to the fact that \(Z^n(\cdot) \leq n\). Again, by the definition of relative compact set in \(M\), we have that \(\sup_n \bar{R}^n(0)(\mathbb{R}) = M_0 < \infty\). It follows from the dynamic equation \((5.2)\) and \((5.3)\) that for all \(x > 0\),

\[
\bar{R}^n(t)(C_x) \leq \bar{R}^n(0)(C_x) + \frac{1}{n} \sum_{i=1}^{E^n(t)} \delta_{u^n_i}(C_x),
\]

\[
\bar{Z}^n(t)(C_x) \leq \bar{Z}^n(0)(C_x) + \frac{1}{n} \sum_{i=1}^{E^n(t)} \delta_{v^n_i}(C_x).
\]

Denote \(\tilde{L}^n_1(t) = \frac{1}{n} \sum_{i=1}^{E^n(t)} \delta_{u^n_i}\) and \(\tilde{L}^n_2(t) = \frac{1}{n} \sum_{i=1}^{E^n(t)} \delta_{v^n_i}\). Let us first study these two terms. Recall the definition of the event \(\Omega^n_{GC}(M, L)\) and the envelope function \(\bar{f}\) (which increases to infinity) in (B.7). For the application here, it is enough to set \(M = 1\) and \(L = 2\lambda T\). On the event \(\Omega^n_E \cap \Omega^n_{GC}(M, L)\), we have

\[
\langle \bar{f}, \tilde{L}^n_1(t) \rangle \leq \langle \bar{f}, \frac{2\lambda T}{n} \sum_{i=1}^{2\lambda T n} \delta_{u^n_i} \rangle \leq 2\lambda T \langle \bar{f}, \nu_F \rangle + 1,
\]

for all large enough \(n\). Similarly, on the same event we have that

\[
\langle \bar{f}, \tilde{L}^n_2(t) \rangle \leq \langle \bar{f}, \frac{2\lambda T}{n} \sum_{i=1}^{2\lambda T n} \delta_{v^n_i} \rangle \leq 2\lambda T \langle \bar{f}, \nu_G \rangle + 1,
\]
for all large enough $n$. Denote $M_B = 2\lambda T \max(\langle \bar{f}, \nu_F \rangle, \langle \bar{f}, \nu_G \rangle) + 1$. By Markov’s inequality, for all $x > 0$ (again, on the same event and for all large $n$)

$$
\overline{L}_1^n(t)(C_x) < M_B/\bar{f}(x), \quad \overline{L}_2^n(t)(C_x) < M_B/\bar{f}(x).
$$

Unlike the measure $Z(t) \in M_+$, the measure $\mathcal{R}(t) \in M$. So we need to consider all the test set $C_x = (-\infty, -x)$ for $x \geq 0$. The following inequality again follows from (5.2),

$$
\overline{R}^n(t)(C_x^-) \leq \overline{R}^n(0)(C_x^- + t) + \frac{1}{n} \sum_{i=1}^{E^n(t)} \delta_{u^n_i}(C_x^- + t).
$$

Note that if we take $x > T$, then $\delta_{u^n_i}(C_x^- + t) = 0$. So we have that

$$
\overline{R}^n(t)(C_x^-) \leq \overline{R}^n(0)(C_x^- + T) = \overline{R}^n(0)(C_{x-T}), \quad \text{for all } t \leq T. \tag{5.10}
$$

Now, define the set $K \subset M$ by

$$
K = \left\{ \xi \in M : \xi(\mathbb{R}) < 1 + M_0 + 2\lambda T, \right. \\
\left. \xi(C_x) < \kappa_0(x) + M_B/\bar{f}(x) \text{ for all } x > 0, \right. \\
\left. \xi(C_x^-) \leq \kappa_0(x - T) \text{ for all } x \geq T \right\}.
$$

It is clear that $K$ is relatively compact and on the event $\Omega^n_0 \cap \Omega^G_{GC}(M, L) \cap \Omega^n_0$,

$$
(\overline{R}^n(t), \overline{Z}^n(t)) \in K \times K \text{ for all } t \in [0, T].
$$

The result of this lemma then follows immediately from (5.6), (5.7) and (B.8). \hfill \square

### 5.1.2 Oscillation Bound

The second major step to prove precompactness is to obtain the oscillation bound in Lemma 5.4 below. The oscillation of a càdlàg function $\zeta(\cdot)$ (taking values in a metric space $(\mathbb{E}, \pi)$) on a fixed interval $[0, T]$ is defined as

$$
\omega_T(\zeta(\cdot), \delta) = \sup_{s, t \in [0, T], |s - t| < \delta} \pi[\zeta(s), \zeta(t)].
$$

If the metric space is $\mathbb{R}$, we just use the Euclidean metric; if the space is $M$ or $M_+$, we use the Prohorov metric $d$ defined in Section 1.1. For the measure-valued processes in our model, oscillations mainly result from sudden departures of a large number of customers. To control the departure process, we show that $\overline{Z}^n(\cdot)$ and $\overline{R}^n(\cdot)$ assign arbitrarily small mass to small intervals.

**Lemma 5.3.** Assume (3.10), (3.18)–(3.21). Fix $T > 0$. For each $\epsilon, \eta > 0$ there exists a $\kappa > 0$ (depending on $\epsilon$ and $\eta$) such that

$$
\liminf_{n \to \infty} \mathbb{P}^n\left( \sup_{t \in [0, T]} \ \sup_{x \in \mathbb{R}_+} \overline{R}^n(t)([x, x + \kappa]) \leq \epsilon \right) \geq 1 - \eta. \tag{5.11}
$$

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Proof. First, we have that for any $\epsilon, \eta > 0$, there exists a $\kappa$ such that

$$\liminf_{n \to \infty} \mathbb{P}^n \left( \sup_{x \in \mathbb{R}_+} \tilde{Z}^n(0)([x, x + \kappa]) \leq \epsilon / 2 \right) \geq 1 - \eta.$$  \hspace{1cm} (5.12)

This inequality is derived from the initial condition. The derivation is exactly the same as in the proof of (5.14) in [31], so we omit it here for brevity.

Now we need to extend this result to the interval $[0, T]$. Denote the event in (5.12) by $\Omega^0_n$, and the event in Lemma 5.1 by $\Omega^c_n(K)$. Fix $M = 1$ and $L = 2\lambda T$, let

$$\Omega^1_n(M, L) = \Omega^0_n \cap \Omega^c_n(K) \cap \Omega^n_E \cap \Omega^n_{GC}(M, L).$$  \hspace{1cm} (5.13)

By (5.12), Lemma 5.1, (5.6) and (B.8), for any fixed $M, L > 0$,

$$\liminf_{n \to \infty} \mathbb{P}^n \left( \Omega^1_n(M, L) \right) \geq 1 - \eta.$$

In the remainder of the proof, all random objects are evaluated at a fixed sample path in $\Omega^1_n(M, L)$.

It follows from the fluid scaled stochastic dynamic equation (5.3) that

$$\tilde{Z}^n(t)([x, x + \kappa]) \leq \tilde{Z}^n(0)([x, x + \kappa] + t)$$

$$+ \frac{1}{n} \sum_{i = B(t) + 1}^{B(t)} \delta_{v^n_i}([x, x + \kappa] + t - \tau^n_i),$$

for each $x, \kappa \in \mathbb{R}_+$. By (5.12), the first term on the right hand side of the above equation is always upper bounded by $\epsilon / 2$. Let $S$ denote the second term on the right hand side of the preceding equation. Now it only remains to show that $S < \epsilon / 2$.

Let $0 = t_0 < t_1 < \cdots < t_J = t$ be a partition of the interval $[0, t]$ such that $|t_{j+1} - t_j| < \delta$ for all $j = 0, \cdots, J - 1$, where $\delta$ and $N$ are to be chosen below. Write $S$ as the summation

$$S = \sum_{j=0}^{J-1} \frac{1}{n} \sum_{i=B(t_j)+1}^{B(t_{j+1})} \delta_{v^n_i}([x, x + \kappa] + t - \tau^n_i).$$

Recall that $\tau^n_i$ is the time that the $i$th job starts service, so on each sub-interval $[t_j, t_{j+1}]$ those $i$’s to be summed must satisfy $t_j \leq \tau^n_i \leq t_{j+1}$. This implies that

$$t - t_{j+1} \leq t - \tau_i \leq t - t_j.$$

Then

$$S \leq \sum_{j=0}^{J-1} \frac{1}{n} \sum_{i=B(t_j)+1}^{B(t_{j+1})} \delta_{v^n_i}([x + t - t_{j+1}, x + t - t_j + \kappa]).$$

By (5.1), we have for all $j = 0, \cdots, J - 1$

$$-\tilde{R}^n(0) \leq \tilde{B}^n(t_j) \leq \tilde{E}^n(T),$$

$$0 \leq \tilde{B}^n(t_{j+1}) - \tilde{B}^n(t_j) \leq \tilde{E}^n(T) + \tilde{R}^n(0).$$
By Lemmas 5.1 and 5.2, \( \bar{R}^n(0) < M_0 \) and \( \bar{E}^n(T) \leq 2\lambda T \) on \( \Omega^n_E(K) \cap \Omega^n_E \) for some constant \( M_0 \). Take \( M = \max(M_0, 2\lambda T) \) and \( L = M_0 + 2\lambda T \), it follows from the Glevenko-Cantelli estimate (B.7) that

\[
\frac{1}{n} \sum_{i=B^n(t_j)+1}^{B^n(t_{j+1})} \delta_{\nu^n}([x + t - t_j + 1, x + t - t_j + \kappa]) \leq (\bar{B}^n(t_{j+1}) - \bar{B}^n(t_j)) \nu^n([x + t - t_j + 1, x + t - t_j + \kappa]) + \frac{\epsilon}{4J},
\]

for each \( j < J \). By condition (3.19), for any \( \epsilon_2 > 0 \),

\[
d[\nu^n_G, \nu_G] < \epsilon_2,
\]

for all large \( n \). By the definition of Prohorov metric, we have

\[
\nu^n_G([x + t - t_j + 1, x + t - t_j + \kappa]) \leq \nu_G([x + t - t_j + 1 - \epsilon_2, x + t - t_j + \kappa + \epsilon_2]),
\]

for all large \( n \). Since \( [x + t - t_j + 1 - \epsilon_2, x + t - t_j + \kappa + \epsilon_2] \) is a close interval with length less than \( \kappa + \delta + 2\epsilon_2 \), by condition (3.10), we can choose \( \kappa, \delta, \epsilon_2 \) small enough such that

\[
\nu([x + t - t_j + 1 - \epsilon_2, x + t - t_j + \kappa + \epsilon_2]) \leq \frac{\epsilon}{4M}.
\]

Thus, we conclude that

\[
S \leq \frac{\epsilon}{4J}[\bar{B}^n(T) - \bar{B}^n(0)] + \frac{\epsilon}{4} \leq \epsilon/2.
\]

This completes the proof.

**Lemma 5.4.** Assume (3.10), (3.18)–(3.21). Fix \( T > 0 \). For each \( \epsilon, \eta > 0 \) there exists a \( \delta > 0 \) (depending on \( \epsilon \) and \( \eta \)) such that

\[
\liminf_{n \to \infty} \mathbb{P}^n \left( \mathbb{w}_T((\bar{R}^n, \bar{Z}^n)(\cdot), \delta) \leq 3\epsilon \right) \geq 1 - \eta. \tag{5.14}
\]

**Proof.** Define

\[
\Omega^n_{\text{Reg}}(\epsilon, \kappa) = \left\{ \sup_{t \in [0, T]} \sup_{x \in \mathbb{R}_+} \bar{Z}^n(t)([x, x + \kappa]) \leq \epsilon \right\}.
\]

By (5.6) and Lemma 5.3, for each \( \epsilon, \eta > 0 \) there exists a \( \kappa > 0 \) such that

\[
\liminf_{n \to \infty} \mathbb{P}^n \left( \Omega^n_E \cap \Omega^n_{\text{Reg}}(\epsilon, \kappa) \right) > 1 - \eta. \tag{5.15}
\]

On the event \( \Omega^n_E \cap \Omega^n_{\text{Reg}}(\epsilon, \kappa) \), we have some control over the dynamics of the system. First, note that the number of customers (in the virtual buffer, including those who have abandoned but ought to get service if they did not) that enter the server during time interval \((s, t]\) can be upper bounded by

\[
\bar{B}^n(s, t) \leq \bar{E}^n(s, t) + \bar{Z}^n(s)([0, t - s]).
\]
When \( t - s \leq \min(\frac{\epsilon}{2\lambda}, \kappa) \), by the definition of \( \Omega^n_E \) and \( \Omega^n_{\text{Reg}}(\epsilon, \kappa) \), we have
\[
\bar{E}^n(s, t) \leq \epsilon \quad \text{(5.16)} \\
\bar{B}^n(s, t) \leq 2\epsilon. \quad \text{(5.17)}
\]

Second, by the dynamic equation (5.2), for any \( s < t \) and any set \( C \in \mathcal{B}(\mathbb{R}) \),
\[
\bar{R}^n(t)(C) - \bar{R}^n(s)(C^{3\epsilon})) \leq \bar{B}^n(s, t) + \bar{E}^n(s, t) \\
+ \frac{1}{n} \sum_{1+B^n(t)} E^n(s) \left[ \delta \mu^n_i(C + t - a^n_i) - \delta \mu^n_i(C^{3\epsilon} + s - a^n_i) \right],
\]
where \( C^a \) is the \( a \)-enlargement of the set \( C \) as defined in Section 1.1. Note that when \( t - s \leq 3\epsilon \), \( C + t - a^n_i \subseteq C^{3\epsilon} + s - a^n_i \) for all \( i \in \mathbb{Z} \), which implies that the second term in the above inequality is less than zero. By (5.16) and (5.17),
\[
\bar{R}^n(t)(C) - \bar{R}^n(s)(C^{3\epsilon})) \leq 3\epsilon.
\]

By Property (ii) on page 72 in [1], we have
\[
d[\bar{R}^n(t), \bar{R}^n(s)] \leq 3\epsilon. \quad \text{(5.18)}
\]

Finally, by the dynamic equation (5.3),
\[
\bar{Z}^n(t)(C) \leq \bar{Z}^n(s)(C + t - s) + \bar{B}^n(s, t).
\]

Note that when \( t - s \leq 2\epsilon \), \( C + t - s \subseteq C^{2\epsilon} \), where \( C^a \) is the \( a \)-enlargement of the set \( C \) as defined in Section 1.1. By (5.17), we have
\[
\bar{Z}^n(t)(C) \leq \bar{Z}^n(s)(C^{2\epsilon}) + 2\epsilon.
\]

By Property (ii) on page 72 in [1], we have
\[
d[\bar{Z}^n(s), \bar{Z}^n(t)] \leq 2\epsilon. \quad \text{(5.19)}
\]

The result of this lemma follows immediately from (5.15), (5.18) and (5.19).

\[\square\]

### 5.2 Convergence to the Fluid Model Solution

We have established the precompactness in Theorem 5.1. So every subsequence of the fluid scaled processes has a further subsequence which converges to some limit. For simplicity of notations, we index the convergent subsequence again by \( n \). So we have that
\[
(\bar{R}^n(\cdot), \bar{Z}^n(\cdot)) \Rightarrow (\bar{R}(\cdot), \bar{Z}(\cdot)) \quad \text{as} \ n \to \infty. \quad \text{(5.20)}
\]

By the oscillation bound in Lemma 5.4, the limit \((\bar{R}(\cdot), \bar{Z}(\cdot))\) is almost surely continuous. We have the following result that further characterizes the above limit.
Lemma 5.5. Assume (3.10)–(3.13) and (3.18)–(3.21). The limit \((\mathcal{R}(\cdot), \mathcal{Z}(\cdot))\) in (5.20) is almost surely the solution to the fluid model \((\lambda, F, G)\) with initial condition \((\mathcal{R}_0, \mathcal{Z}_0)\).

The rest of this section is devoted to characterizing the limits. To better structure the proof, we first provide some preliminary estimates based on the dynamic equations (5.2) and (5.3).

Lemma 5.6. Let \(\{t_j\}_{j=0}^J\) be a partition of the interval \([s, t]\) such that \(s = t_0 < t_1 < \ldots < t_J = t\). We have for any \(x \in \mathbb{R}\),

\[
\mathcal{R}^n(t)(C_x) \leq \sum_{i=0}^{J-1} \frac{1}{n} \sum_{i=1+E^n(t_j)}^{E^n(t_{j+1})} \delta_{u_i^n}(C_x + t - t_j) + |E^n(s) - \bar{B}^n(t)|, \tag{5.21}
\]

\[
\mathcal{R}^n(t)(C_x) \geq \sum_{i=0}^{J-1} \frac{1}{n} \sum_{i=1+E^n(t_j)}^{E^n(t_{j+1})} \delta_{u_i^n}(C_x + t - t_j) - |E^n(s) - \bar{B}^n(t)|. \tag{5.22}
\]

If in addition that \(\sup_{\tau \in [s,t]} |\bar{E}^n(\tau) - \lambda \tau| < \epsilon\), then for any \(x > 0\),

\[
\mathcal{Z}^n(t)(C_x) \leq \mathcal{Z}^n(s)(C_x + t - s) + \sum_{j=0}^{J-1} \frac{1}{n} \sum_{i=1+B^n(t_j)}^{B^n(t_{j+1})} \delta_{\bar{u}_i^n}(C_0 + \frac{\bar{R}^n_{L,j} - 2\epsilon}{\lambda})\delta_{\bar{u}_i^n}(C_x + t - t_j), \tag{5.23}
\]

\[
\mathcal{Z}^n(t)(C_x) \geq \mathcal{Z}^n(s)(C_x + t - s) + \sum_{j=0}^{J-1} \frac{1}{n} \sum_{i=1+B^n(t_j)}^{B^n(t_{j+1})} \delta_{\bar{u}_i^n}(C_0 + \frac{\bar{R}^n_{U,j} + 2\epsilon}{\lambda})\delta_{\bar{u}_i^n}(C_x + t - t_{j+1}), \tag{5.24}
\]

where \(\bar{R}^n_{L,j} = \inf_{t \in [t_j, t_{j+1}]} \bar{R}^n(t)\) and \(\bar{R}^n_{U,j} = \sup_{t \in [t_j, t_{j+1}]} \bar{R}^n(t)\).

Proof. Note that \(0 \leq \delta_{u^n_i}(C) \leq 1\) for any Borel set \(C\) and any random variable \(u^n_i\). So by the dynamic equation (5.2), we have

\[
\left| \mathcal{R}^n(t)(C) - \frac{1}{n} \sum_{i=E^n(s)+1}^{E^n(t)} \delta_{u^n_i}(C + t - a^n_i) \right| \leq |\bar{E}^n(s) - \bar{B}^n(t)|.
\]

For those \(i\)'s such that \(E^n(t_j) < i \leq E^n(t_{j+1})\), we have that

\[
t_j < a^n_i \leq t_{j+1}. \tag{5.25}
\]

This implies that \(C_x + t - a_i \subseteq C_x + t - t_j\). So we have

\[
\sum_{i=1+E^n(t_j)}^{E^n(t_{j+1})} \delta_{u^n_i}(C_x + t - a_i) \leq \sum_{i=1+E^n(t_j)}^{E^n(t_{j+1})} \delta_{u^n_i}(C_x + t - t_j).
\]

This establishes (5.21). Also, (5.25) implies \(C_x + t - t_{j+1} \subseteq C_x + t - a_i\). So (5.22) follows in the same way.
For those $i$'s such that $B^n(t_j) < i \leq B^n(t_{j+1})$, we have that

$$t_j < \tau_i^n \leq t_{j+1}.$$ 

Note that $\bar{R}^n(\tau_i^n) = \bar{E}^n(\tau_i^n) - \bar{E}^n(a^n_i)$ for each $i$. So, by the closeness between $\bar{E}^n(\cdot)$ and $\lambda$, we have

$$|\bar{R}^n(\tau_i^n) - \lambda(\tau_i^n - a^n_i)|$$

$$\leq |\bar{R}^n(\tau_i^n) - \bar{E}^n(\tau_i^n) + \bar{E}^n(a^n_i)| + |\bar{E}^n(\tau_i^n) - \bar{E}^n(a^n_i) - \lambda(\tau_i^n - a^n_i)|$$

$$\leq 2\epsilon.$$

So

$$\bar{R}^n_{L,j} - 2\epsilon \leq \lambda(\tau_i^n - a^n_i) \leq \bar{R}^n_{U,j} + 2\epsilon,$$

for all $i$'s such that $B^n(t_j) < i \leq B^n(t_{j+1})$. Thus,

$$\sum_{i=1+B^n(t_j)}^{B^n(t_{j+1})} \delta_\tau^n(C_0 + \tau_i^n - a^n_i)\delta_\tau^n(C_x + t - \tau_j^n) \leq \sum_{i=1+B^n(t_j)}^{B^n(t_{j+1})} \delta_\tau^n(C_0 + \frac{\bar{R}^n_{L,j} - 2\epsilon}{\lambda})\delta_\tau^n(C_x + t - t_j).$$

This implies (5.23). And (5.24) can be proved in the same way. □

Recall the notations $\tilde{E}^n(m, l), \tilde{L}^n_0(m, l)$ and $\tilde{L}^n_0(m, l)$ are defined in (B.1)–(B.3) in the appendix. Using these notations, Lemma 5.6 can be written as the following:

**Lemma 5.7.** Let $\{t_j\}_{j=0}^J$ be a partition of the interval $[s, t]$ such that $s = t_0 < t_1 < \ldots < t_J = t$. We have for any $x \in \mathbb{R}$,

$$\bar{R}^n(t)(C_x) \leq \sum_{i=0}^{J-1} 1_{(C_x + t - t_j)}\tilde{L}^n_p(E^n(t_j), \bar{E}^n(t_j, t_{j+1})) + |\bar{E}^n(s) - \bar{B}^n(t)|, \quad (5.26)$$

$$\tilde{R}^n(t)(C_x) \geq \sum_{i=0}^{J-1} 1_{(C_x + t - t_{j+1})}\tilde{L}^n_p(E^n(t_j), \bar{E}^n(t_j, t_{j+1})) - |\bar{E}^n(s) - \bar{B}^n(t)|. \quad (5.27)$$

If in addition that $\sup_{\tau \in [s, t]}|\bar{E}^n(\tau) - \lambda\tau| < \epsilon$, then for any $x > 0$,

$$\bar{Z}^n(t)(C_x) \leq \bar{Z}^n(s)(C_x + t - s)$$

$$+ \sum_{j=0}^{J-1} 1_{(C_0 + \frac{\bar{R}^n_{L,j} - 2\epsilon}{\lambda} \times (C_x + t - t_j)}\tilde{L}^n(B^n(t_j), \bar{B}^n(t_j, t_{j+1})), \quad (5.28)$$

$$\tilde{Z}^n(t)(C_x) \geq \tilde{Z}^n(s)(C_x + t - s)$$

$$+ \sum_{j=0}^{J-1} 1_{(C_0 + \frac{\bar{R}^n_{L,j} - 2\epsilon}{\lambda} \times (C_x + t - t_{j+1})\tilde{L}^n(B^n(t_j), \bar{B}^n(t_j, t_{j+1})). \quad (5.29)$$

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Fix a constant $T > 0$ and let $M = 1$ and $L = 2\lambda T$. Denote the random variable

$$
\hat{V}_{M,L}^n = \max_{-nM < m < nM} \sup_{t \in [0,L]} \sup_{x \in \mathbb{R}} \left\{ \left| \tilde{L}^n(m,l)(C_x \times C_y) - \nu_F^n(C_x)\nu_G^n(C_y) \right| + \left| \tilde{L}^n(m,l)(C_x) - \nu_F^n(C_x) \right| \right\}.
$$

(5.30)

By Lemma B.1, for any fixed constants $M, L > 0$,

$$
\hat{V}_{M,L}^n \Rightarrow 0 \quad \text{as} \quad n \to \infty.
$$

By the assumption (3.18), we have

$$
\hat{E}^n(\cdot) \Rightarrow \lambda \cdot \quad \text{as} \quad n \to \infty.
$$

Since both the above two limits are deterministic, those convergences are joint with the convergence of $(\tilde{R}^n(\cdot), \tilde{Z}^n(\cdot))$. Now, for each $n \geq 1$, we can view $(\hat{E}^n(\cdot), \hat{R}^n(\cdot), \hat{Z}^n(\cdot), V_{M,L})$ as a random variable in the space $E_1$, which is the product space of three $D([0,\infty), \mathbb{R})$ spaces and the space $\mathbb{R}$. And $(\tilde{L}(m,\cdot), \tilde{L}_n^p(m,\cdot), \tilde{L}_n^o(m,\cdot) : m \in \mathbb{Z})$ in the product space $E_2$ of countable many $D([0,\infty), M)$ spaces. It is clear that both $E_1$ and $E_2$ are complete and separable metric spaces. Using the extension of Skorohod representation Theorem, Lemma C.1, we assume without loss of generality that $E^n(\cdot), \tilde{R}^n(\cdot), \tilde{Z}^n(\cdot), \tilde{V}_{M,L}^n, L^n(m,\cdot), L_n^p(m,\cdot), L_n^o(m,\cdot), m \in \mathbb{Z}$, and $(\tilde{R}(\cdot), \tilde{Z}(\cdot))$ are defined on a common probability space $(\Omega, \mathbb{F}, \mathbb{P})$ such that, almost surely,

$$
\left( (\tilde{R}^n(\cdot), \tilde{Z}^n(\cdot)), \tilde{V}_{M,L}^n, \tilde{E}^n(\cdot) \right) \Rightarrow \left( (\tilde{R}(\cdot), \tilde{Z}(\cdot)), 0, \lambda \cdot \right) \quad \text{as} \quad n \to \infty,
$$

(5.31)

and inequalities (5.26)--(5.29) and equation (5.30) also hold almost surely. Note that the convergence of each function component in the above is in the Skorohod $J_1$ topology. Since the limit is continuous, the convergence is equivalent to the convergence in the uniform norm on compact intervals. Thus as $n \to \infty$,

$$
\sup_{t \in [0,T]} d[\tilde{R}^n(t), \tilde{R}(t)] \to 0,
$$

(5.32)

$$
\sup_{t \in [0,T]} d[\tilde{Z}^n(t), \tilde{Z}(t)] \to 0,
$$

(5.33)

$$
\sup_{t \in [0,T]} |\tilde{E}^n(t) - \lambda t| \to 0,
$$

(5.34)

where $d$ is the Skorohod metric defined in Section 1.1. Same as on the original probability space, let

$$
\tilde{R}^n(\cdot) = \langle 1, \hat{R}^n(\cdot) \rangle, \quad \tilde{Q}^n(\cdot) = \langle 1_{(0,\infty)}, \hat{R}^n(\cdot) \rangle,
$$

$$
\tilde{Z}^n(\cdot) = \langle 1, \hat{Z}^n(\cdot) \rangle, \quad \tilde{X}^n(\cdot) = \hat{Q}^n(\cdot) + \tilde{Z}^n(\cdot),
$$

and

$$
\tilde{B}^n(\cdot) = \tilde{E}^n(\cdot) - \tilde{R}^n(\cdot).
$$
According to (5.32) and (5.34), we have
\[
\sup_{t \in [0,T]} |\tilde{B}^n(t) - \tilde{B}(t)| \to 0. \tag{5.35}
\]

For each \(n\), let \(\tilde{\Omega}_{n,2}\) be an event of probability one on which the stochastic dynamic equations (5.2) and (5.3) and the policy constraints (2.6) and (2.7) hold. Define \(\tilde{\Omega}_0 = \tilde{\Omega}_1 \cap (\cap_{n=0}^{\infty} \tilde{\Omega}_{n,2})\), where \(\tilde{\Omega}_1\) is the event of probability one on which (5.31) holds. Then \(\tilde{\Omega}_0\) also has probability one. Based on Lemma 5.6 and the above argument using Skorohod Representation theorem, we can now prove Lemma 5.5.

**Proof of Lemma 5.5.** For any \(t \geq 0\), fix a constant \(T > t\). Let us now study \((\tilde{R}(\cdot), \tilde{Z}(\cdot))\) on the time interval \([0,T]\). It is enough to show that on the event \(\tilde{\Omega}_0\), \((\tilde{R}(t), \tilde{Z}(t))\) satisfies the fluid model equation (3.1)-(3.2) and the constraints (3.3)-(3.4). Assume for the remainder of this proof that all random objects are evaluated at a sample path in the event \(\tilde{\Omega}_0\).

We first verify (3.1). For any \(\epsilon > 0\), consider the difference
\[
\tilde{R}(t)(C_x) - \tilde{R}(t)(0) F^c(x + t - s) d\lambda s
\]
\[
= \tilde{R}(t)(C_x) - \tilde{R}^n(t)(C_x) + \tilde{R}^n(t)(C_x) - \tilde{R}(t)(0) F^c(x + t - s) d\lambda s,
\]
where \(C_x^\epsilon\) is the \(\epsilon\)-enlargement of the set \(C_x\) as defined in Section 1.1, which is essentially \(C_{x-\epsilon}\). Let \(t_0 = t - \tilde{R}(t)/\lambda\). According to (5.26), we have that
\[
\tilde{R}(t)(C_x) - \tilde{R}(t)(0) F^c(x + t - s) d\lambda s
\]
\[
\leq \tilde{R}(t)(C_x) - \tilde{R}^n(t)(C_x^\epsilon) + \|\mathcal{E}^n(t_0) - \tilde{B}^n(t)\| \tag{5.36}
\]
\[
\sum_{j=0}^{J-1} \{1(C_x^\epsilon + t - t_j), \mathcal{E}_p^w(E^n(t_j), E^n(t_j, t_{j+1})) - \int_{t_0}^t F^c(x + t - s) d\lambda s,
\]
where \(\{t_j\}_{j=0}^J\) is a partition of the interval \([t_0, t]\) such that \(t_0 < t_1 < \ldots < t_J = t\) and \(\max_j (t_{j+1} - t_j) < \delta\) for some \(\delta > 0\). By the definition of Prohorov metric and the convergence in (5.32), the first term on the right hand side of (5.36) is bounded by \(\epsilon\) for all large \(n\). By (5.32) and (5.34)
\[
|\tilde{B}^n(t) - \mathcal{E}^n(t_0)| = |\mathcal{E}^n(t) - \tilde{R}^n(t) - \tilde{B}^n(t_0)|
\]
\[
\leq |\mathcal{E}^n(t) - \lambda t| + |\tilde{R}^n(t) - \tilde{R}(t)| + |\mathcal{E}^n(t_0) - \lambda t_0| < 3\epsilon,
\]
for all large \(n\). So
\[
\tilde{R}(t)(C_x) - \tilde{R}(t)(0) F^c(x + t - s) d\lambda s
\]
\[
\leq 4\epsilon + \sum_{j=0}^{J-1} \{1(C_x^\epsilon + t - t_j), \mathcal{E}_p^w(E^n(t_j), E^n(t_j, t_{j+1})) - \int_{t_0}^t F^c(x + t - s) d\lambda s, \tag{5.37}
\]

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for all large \( n \). Similarly, according to (5.27), we have

\[
\hat{R}(t)(C_x) - \int_{t-\frac{\hat{R}(t)}{\lambda}}^{t} F^c(x + t - s)d\lambda s =
\]

\[
\geq -4\epsilon + \sum_{j=0}^{J-1} \langle 1_{(C_x^j + t - t_j, + 1)}(t_j), \tilde{L}_p^n(E^n(t_j), E^n(t_j, t_{j+1})) \rangle - \int_{t_0}^{t} F^c(x + t - s)d\lambda s,
\]

for all large \( n \). Note that for each \( j \), we have

\[
\langle 1_{(C_x^j + t - t_j, + 1)}(t_j), \tilde{L}_p^n(E^n(t_j), E^n(t_j, t_{j+1})) \rangle
\]

\[
\leq \langle 1_{(C_x^j + t - t_j)}, \tilde{L}_p^n(E^n(t_j), \lambda(t_{j+1} - t_j) + 2\epsilon) \rangle
\]

\[
\leq \left[ \lambda(t_{j+1} - t_j) + 2\epsilon \right] \nu^p_\lambda(C_x^j + t - t_j) + \epsilon
\]

\[
\leq \left[ \lambda(t_{j+1} - t_j) + 2\epsilon \right] \nu_F(C_x + t - t_j) + \epsilon + \epsilon + \epsilon
\]

\[
\leq \lambda(t_{j+1} - t_j) \nu_F(C_x + t - t_j) + (3 + \lambda \delta)\epsilon
\]

for all large \( n \), where the first inequality is due to (5.34), the second one is due to (5.31) (the component of \( V_{\hat{M},L} \)), the third one is due to (3.19), and the last one is due to algebra. Similarly, we can show that

\[
\langle 1_{(C_x^j + t - t_{j+1} - 1)}(t_j), \tilde{L}_p^n(E^n(t_j), E^n(t_j, t_{j+1})) \rangle
\]

\[
\geq \lambda(t_{j+1} - t_j) \nu_F(C_x + t - t_{j+1}) - (3 + \lambda \delta)\epsilon
\]

for all large \( n \). Note that \( \sum_{j=0}^{J-1} \lambda(t_{j+1} - t_j) F^c(x + t - t_j) \) and \( \sum_{j=0}^{J-1} \lambda(t_{j+1} - t_j) F^c(x + t - t_{j+1}) \) serve as the upper and lower Reimann sum of the integral \( \int_{t_0}^{t} F^c(x + t - s)d\lambda s \), which converge to the integration as \( n \rightarrow \infty \). So by (5.37) and (5.38), we have that for all large \( n \),

\[
|\hat{R}(t)(C_x) - \int_{t-\frac{\hat{R}(t)}{\lambda}}^{t} F^c(x + t - s)d\lambda s| \leq (3 + \lambda \delta)J\epsilon + 5\epsilon.
\]

We conclude that \( \hat{R}(t)(C_x) - \int_{t-\frac{\hat{R}(t)}{\lambda}}^{t} F^c(x + t - s)d\lambda s = 0 \) since \( \epsilon \) in the above can be arbitrary. This verifies (3.1).

Next, we verify (3.2). For any \( \epsilon > 0 \), consider the difference

\[
\left| \tilde{Z}(t)(C_x) - \tilde{Z}_0(C_x + t) - \int_{0}^{t} F^c(\frac{\hat{R}(s)}{\lambda}) G^c(x + t - s)d(\lambda s - \hat{R}(s)) \right|
\]

\[
\leq |\tilde{Z}(t)(C_x) - \tilde{Z}_0(C_x + t)| + |\tilde{Z}_0(C_x + t) - \tilde{Z}_0^n(0)(C_x^c + t)|
\]

\[
+ \left| \tilde{Z}_0^n(C_x^c + t) - \tilde{Z}_0^n(0)(C_x^c + t) - \int_{0}^{t} F^c(\frac{\hat{R}(s)}{\lambda}) G^c(x + t - s)d(\lambda s - \hat{R}(s)) \right|
\]

where the above inequality is due to the fluid scaled stochastic dynamic equation (5.3). Again, by the definition of Prohorov metric and the convergence in (5.33), each of the first two terms on the
right hand side in the above inequality is less than \( \epsilon \) for all large \( n \). Let \( \{t_j\}_{j=0}^J \) be a partition of the interval \([0, t]\) such that \( 0 = t_0 < t_1 < \ldots < t_J = t \) and \( \max_j (t_{j+1} - t_j) < \delta \) for some \( \delta > 0 \). Let

\[
\tilde{R}_{U,j} = \sup_{t \in [t_j, t_{j+1}]} \tilde{R}(t), \quad \tilde{R}_{L,j} = \inf_{t \in [t_j, t_{j+1}]} \tilde{R}(t).
\]

By (5.32), we have that

\[
|\tilde{R}_{U,j}^n - \tilde{R}_{U,j}| \leq \epsilon, \quad |\tilde{R}_{L,j}^n - \tilde{R}_{L,j}| \leq \epsilon,
\]

for all large \( n \). So for each \( j \), we have

\[
\langle 1_{(C_0 + \frac{R_{L,j} - 3\epsilon}{\lambda}) \times (C_0 + t - t_j)} \rangle \mathcal{L}^n(B^n(t_j), \tilde{B}^n(t_j, t_{j+1})) \\
\leq \langle 1_{(C_0 + \frac{R_{L,j} - 3\epsilon}{\lambda}) \times (C_0 + t - t_j)} \rangle \mathcal{L}^n(B^n(t_j), \tilde{B}(t_{j+1}) - \tilde{B}(t_j) + 2\epsilon) \\
\leq [\tilde{B}(t_{j+1}) - \tilde{B}(t_j) + 2\epsilon] \mu_F(C_0 + \frac{R_{L,j} - 3\epsilon}{\lambda})\nu_G(C_x + t - t_j) + \epsilon \\
\leq [\tilde{B}(t_{j+1}) - \tilde{B}(t_j) + 2\epsilon] [\mu_F(C_0 + \frac{R_{L,j} - 3\epsilon}{\lambda}) + \frac{3\epsilon}{\lambda}] [\mu_G(C_x + t - t_j) + \epsilon + \epsilon]
\]

for all large \( n \), where the first inequality is due to (5.35), the second one is due to (5.31) (the component of \( V_{M,L}^\alpha \)), the third one is due to (3.19). Let \( M_B \) be a finite upper bound of \( \tilde{B}(t_j) - \tilde{B}(t_0) \), the above inequality can be further bounded by

\[
[\tilde{B}(t_{j+1}) - \tilde{B}(t_j)] \mu_F(C_0 + \frac{R_{L,j} - 3\epsilon}{\lambda})\nu_G(C_x + t - t_j) + \left( \frac{3}{\lambda} + 2 \right) M_B \epsilon + 3\epsilon.
\]

Similarly, we can show that

\[
\langle 1_{(C_0 + \frac{R_{L,j} + 2\epsilon}{\lambda}) \times (C_0 + t - t_{j+1})} \rangle \mathcal{L}^n(B^n(t_j), \tilde{B}^n(t_j, t_{j+1})) \\
\geq [\tilde{B}(t_{j+1}) - \tilde{B}(t_j)] \mu_F(C_0 + \frac{R_{L,j}}{\lambda})\nu_G(C_x + t - t_j) - \left( \frac{3}{\lambda} + 2 \right) M_B \epsilon - 3\epsilon.
\]

Note that \( \sum_{j=1}^{J-1} [\tilde{B}(t_{j+1}) - \tilde{B}(t_j)] F_c(\tilde{R}_{L,j}) G_c(x + t - t_j) \) and \( \sum_{j=0}^{J-1} [\tilde{B}(t_{j+1}) - \tilde{B}(t_j)] F_c(\tilde{R}_{L,j}) G_c(x + t - t_{j+1}) \) serve as the upper and lower Reimann sum of the integral \( \int_0^t F_c(\tilde{R}(s)) G_c(x + t - s) d\tilde{B}(s) \), which converge to the integration as \( n \to \infty \). So, by (5.28) and (5.29), we have that for all large \( n \),

\[
\left| \tilde{Z}^n(t)(C_x) - \tilde{Z}^n(0)(C_x) + t - \int_0^t F_c(\tilde{R}(s)) G_c(x + t - s) d\tilde{B}(s) \right| \leq \left( \frac{3}{\lambda} + 2 \right) M_B \epsilon + 3\epsilon + \epsilon.
\]

In summary, the right hand side of (5.39) can be bounded by a finite multiple of \( \epsilon \). We conclude that the left hand side of (5.39) must be 0 since it does not depend on \( \epsilon \), which can be arbitrary. This verifies (3.2).

The verification of fluid constrains (3.3) and (3.4) is quite straightforward. Basically, it is just

\[
\tilde{Q}^n(t) = (\tilde{X}^n(t) - 1)^+, \\
\tilde{Z}^n(t) = (\tilde{X}^n(t) \wedge 1),
\]

to \( n \to \infty \). We omit it for brevity. \( \square \)
6 The Special Case with Exponential Distribution

In this section, we verify that the fluid model developed in this paper for the general patience and service time distributions is consistent with the one in [27], that was obtained in the special case where both distributions are assumed to be exponential.

Our fluid model equations implies the key relationship (4.5). Now, we specialize in the case with exponential distribution, i.e.

\[ F(t) = F_e(t) = 1 - e^{-\alpha t}, \quad G(t) = G_e(t) = 1 - e^{-\mu t}, \quad \text{for all } t \geq 0. \]

Now (4.5) becomes

\[ \bar{X}(t) = \zeta_0(t) + \rho \int_0^t \left[ 1 - \frac{\alpha}{\lambda}((\bar{X}(t - s) - 1)^+) \right] \mu e^{-\mu s} ds + \int_0^t (\bar{X}(t - s) - 1)^+ \mu e^{-\mu s} ds. \]

In the case of exponential service time distribution, the remaining service time of those initially in service and the service times of those initially waiting in queue are also assumed to be exponentially distributed. So we have

\[ \zeta_0(t) = \bar{Z}_0(C_0 + t) + \bar{Q}_0 e^{-\mu t} = \bar{X}_0 e^{-\mu t}, \]

where \( \bar{X}_0 = \bar{Z}_0 + \bar{Q}_0 \) is the initial number of customers in the system. By some algebra, the above two equations can be simplified as the following,

\[ \bar{X}(t) = \bar{X}_0 e^{-\mu t} + \rho [1 - e^{-\mu t}] + (\mu - \alpha) \int_0^t (\bar{X}(t - s) - 1)^+ e^{-\mu s} ds. \]  (6.1)

By the change of variable \( t - s \to s \), the above integration can be written as

\[ \int_0^t (\bar{X}(t - s) - 1)^+ e^{-\mu s} ds = e^{-\mu t} \int_0^t (\bar{X}(s) - 1)^+ e^{\mu s} ds. \]

Taking the derivative on both sides of (6.1) yields

\[ \bar{X}'(t) = -\mu \bar{X}_0 e^{-\mu t} + \mu \rho e^{\mu t} + (\mu - \alpha) \bar{X}(t - 1)^+ e^{-\mu t} \]

\[ + (\mu - \alpha) \left[ -\mu e^{-\mu t} \int_0^t (\bar{X}(s) - 1)^+ e^{\mu s} ds + e^{-\mu t} (\bar{X}(t) - 1)^+ e^{\mu t} \right] \]

\[ = -\mu \bar{X}_0 e^{-\mu t} - \mu \rho [1 - e^{\mu t}] + \mu \rho + (\mu - \alpha) \bar{X}(t - 1)^+ \]

\[ = -\mu \bar{X}(t) + \mu \rho + (\mu - \alpha)(\bar{X}(t) - 1)^+. \]

Using the notation in [27], \( a^- = -\min(0, a) \) for any \( a \in \mathbb{R} \). Note that \( a = \min(a, 1) + (a - 1)^+ = 1 - (a - 1)^- + (a - 1)^+ \). So the above equation further implies

\[ \bar{X}'(t) = \mu (\rho - 1) + \alpha (\bar{X}(t) - 1)^+ + \mu (\bar{X}(t) - 1)^-, \quad \text{for all } t \geq 0. \]

This equation is consistent with Theorem 2.2 in [27] (\( \mu \) is assumed to be 1 in that paper).
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### A Convolution Equation

**Lemma A.1.** Assume that $G(\cdot)$ is a distribution function with $G(0) < 1$, $\zeta(\cdot) \in D([0, T], \mathbb{R})$, $H(\cdot)$ is a Lipschitz continuous function, and $\rho \in \mathbb{R}$. There exists a unique solution $x^*(\cdot) \in D([0, T], \mathbb{R})$ to the following equation:

$$
 x(t) = \zeta(t) + \rho \int_0^t H((x(t-s) - 1)^+) dG_e(s) + \int_0^t (x(t-s) - 1)^+ dG(s),
$$

(A.1)

where, $G_e$ is the equilibrium distribution of $G$ as defined in Section 3.1.

**Proof.** Suppose $H(\cdot)$ is Lipschitz continuous with constant $L$. The equilibrium distribution has density $\mu[1-G(\cdot)]$, so $|G_e(t) - G_e(s)| \leq \mu|t-s|$ for any $s, t \in \mathbb{R}$. Since $G(0) < 1$, there exists $b > 0$ such that $\kappa := \rho L [G_e(b) - G_e(0)] + [G(b) - G(0)] < 1$.

Now consider the space $D([0, b], \mathbb{R})$ (all real valued càdlàg functions on $[0, b]$, c.f. Section 1.1) is a subset of the Banach space of bounded, measurable functions on $[0, b]$, equipped with the sup norm. One can check that this subset is closed in the Banach space. Thus, the space $D([0, b], \mathbb{R})$ itself, equipped with the uniform metric $\nu_T$ (defined in Section 1.1), is complete.

For any $y \in D([0, b], \mathbb{R})$, define $\Psi(y)$ by

$$
 \Psi(y)(t) = \zeta(t) + \rho \int_0^t H((y(t-s) - 1)^+) dG_e(s) + \int_0^t (y(t-s) - 1)^+ dG(s),
$$

for any $t \in [0, b]$. By convention, the integration $\int_0^t y(t-s) dF(s)$ is interpreted to be $\int_{[0,t]} y(t-s) dF(s)$ (c.f. Page 43 in [3]). We prove the existence and uniqueness of the solution to equation (A.1)
by showing that \( \Psi \) is a contraction mapping on \( D([0,b], \mathbb{R}) \). According to the proof of Lemma A.1 in [31], the convolution of a càdlàg function with a distribution function is still a càdlàg function. So \( \Psi \) is a mapping from \( D([0,b], \mathbb{R}) \) to \( D([0,b], \mathbb{R}) \). Next, we show that the mapping \( \Psi \) is a contraction. For any \( y, y' \in D([0,b], \mathbb{R}) \), we have that

\[
v_b[\Psi(y), \Psi(y')] \leq \sup_{t \in [0,b]} \rho \int_0^t L[y(t-s) - 1] + (y'(t-s) - 1) dG(s) + \rho L \int_0^b v_b[y, y'] dG(s) + \int_0^b v_b[y, y'] dG(s)
\]

Since \( \kappa < 1 \), the mapping \( \Psi \) is a contraction. By the contraction mapping theorem (c.f. Theorem 3.2 in [12]), \( \Psi \) has a unique fixed point \( x \), i.e. \( x = \psi(x) \). This implies that \( x \in D([0,b], \mathbb{R}) \) is the unique solution to equation (A.1) on \([0,b]\).

It now remains to extend the existence and uniqueness result from \([0,b]\) to \([0,T]\). Denote \( x_b(t) = x(b+t), \zeta_b(t) = \zeta(b+t) + \rho \int_{t}^{b+t} H((x(b+t-s) - 1) dG(s) + \int_{t}^{b+t} (x(b+t-s) - 1) dG(s) \), then we have for \( t \in [0,T-b] \),

\[
x_b(t) = \zeta_b(t) + \rho \int_0^t H((x_b(t-s) - 1) dG(s) + \int_0^t (x_b(t-s) - 1) dG(s). \tag{A.2}
\]

It follows from the previous argument that there is unique solution \( x_b(\cdot) \) to the above equation. Thus, we obtain a unique extension of the solution to (A.1) on the interval \([0,2b]\). Repeating this approach for \( N \) time with \( N \geq [T/b] \) gives a unique solution on the interval \([0,T]\).

\(\Box\)

**Lemma A.2.** Assume the same condition as in Lemma A.1. Let \( x(\cdot) \in D([0,T], \mathbb{R}) \) be the solution to equation (A.1). If \( \rho = \lambda/\mu \) with \( \lambda, \mu > 0 \) (\( \mu \) is the mean of \( G \)), \( H(x) \geq 0 \) for all \( x \geq 0 \), and \( \zeta(\cdot) \) satisfies the following condition

\[
\zeta(t) = h(t) + (\zeta(0) - 1) [1 - G(t)], \tag{A.3}
\]

where \( h(\cdot) \) is a non-increasing function, then the function

\[
(x(t) - 1)^+ - \lambda \int_0^t H((x(s) - 1)^+) \, ds
\]

is non-increasing on the interval \([0,T]\).

**Proof.** To simplify the notation, let \( Q(t) = (x(t) - 1)^+ \) and

\[
D(t) = Q(t) - \lambda \int_0^t H(Q(s)) \, ds \tag{A.4}
\]
for all \( t \in [0, T] \). Since \( G_e(\cdot) \) is the equilibrium distribution, we have

\[
x(t) = \zeta(t) + \rho \int_0^t H(Q(t - s)) \mu[1 - G(s)]ds + \int_0^t Q(t - s)dG(s)
= \zeta(t) + \lambda \int_0^t H(Q(s)) ds - \lambda \int_0^t H(Q(s)) G(t - s)ds + \int_0^t Q(t - s)dG(s).
\]

Applying Fubini’s Theorem (c.f. Theorem 8.4 in [17]) to the second to the last integral in the above, we have

\[
\int_0^t H(Q(s)) G(t - s)ds = \int_0^t \int_0^{t-s} H(Q(s)) dG(\tau)ds
= \int_0^t \int_0^{t-\tau} H(Q(s)) dsdG(\tau).
\]

So we obtain

\[
x(t) - \lambda \int_0^t H(Q(s)) ds = \zeta(t) + \int_0^t \left[ Q(t - s) - \lambda \int_0^{t-s} H(Q(\tau)) d\tau \right] dG(s).
\]

According to the above definition of \( D(\cdot) \), we have

\[
(x(t) \wedge 1) + D(t) = \zeta(t) + \int_0^t D(t - s)dG(s).
\tag{A.5}
\]

It now remains to use (A.5) to show that \( D(\cdot) \) is non-increasing, i.e. for any \( t, t' \in [0, T] \) with \( t \leq t' \), we have \( D(t) \geq D(t') \). Since \( G(0) < 1 \), there exists \( a > 0 \) such that \( G(a) < 1 \). We first show that \( D(\cdot) \) is non-increasing on the interval \([0, a]\). Let

\[
D^* = \sup_{\{(t,t')\in[0,a]\times[0,a]:t\leq t'\}} D(t') - D(t).
\]

Since \( D(\cdot) \) is càdlàg, according to Theorem 6.2.2 in the supplement of [26], it is bounded on the interval \([0, a]\). Thus, \( D^* \) is finite. We will prove by contradiction that \( D^* \leq 0 \), which shows that \( D(\cdot) \) is non-increasing on \([0, a]\). Assume on the contrary that \( D^* > 0 \). Applying (A.5), we have

\[
D(t') - D(t) = (x(t) \wedge 1) - (x(t') \wedge 1) + \zeta(t') - \zeta(t)
+ \int_0^{t'} D(t' - s)dG(s) - \int_0^t D(t - s)dG(s)
= (x(t) \wedge 1) - (x(t') \wedge 1) + \zeta(t') - \zeta(t)
+ \int_t^{t'} D(t' - s)dG(s) + \int_0^t [D(t' - s) - D(t - s)]dG(s).
\]

It follows from (A.1) and (A.4) that \( D(0) = (\zeta(0) - 1) \). This together with condition (A.3) implies that

\[
\zeta(t') - \zeta(t) = h(t') - h(t) + D(0)[G(t) - G(t')].
\tag{A.6}
\]

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If \(x(t') < 1\), by (A.4),
\[
D(t') - D(t) = -\lambda \int_t^{t'} H(Q(s)) \, ds - Q(t),
\]
which is always non-positive; if \(x(t') \geq 1\), then \((x(t) \wedge 1) - (x(t') \wedge 1) \leq 0\). So it follows from (A.7) and \(h(\cdot)\) being non-increasing that
\[
D(t') - D(t) \leq \int_t^{t'} [D(t' - s) - D(0)]dG(s) + \int_0^t [D(t' - s) - D(t - s)]dG(s)
\]
\[
\leq \int_0^{t'} D^*dG(s) = D^*G(t') \leq D^*G(a),
\]
where the last inequality follows from the assumption that \(D^*\) is non-negative. Summarizing both cases of \(x(t')\), we have
\[
D(t') - D(t) \leq \max(0, D^*G(a))
\]
for all \(t, t' \in [0, a] > 0\) with \(t \leq t'\). Taking the supremum on both sides over the set \(\{(t, t') \in [0, a] \times [0, a] : t \leq t'\}\) gives \(D^* \geq F(a)D^*\). This implies that \([1 - G(a)]D^* \leq 0\). Since \(G(a) < 1\), it contradicts the assumption that \(D^* > 0\). So we must have \(D^* \leq 0\), this implies that \(D(\cdot)\) is non-increasing on \([0, a]\). We next extend this property to the interval \([0, T]\) using induction. Suppose we can show that \(D(\cdot)\) is non-decreasing on the interval \([0, na]\) for some \(n \in \mathbb{N}\). Introduce \(D_{na}(t) = D(na + t), x_{na}(t) = x(na + t)\) and
\[
\zeta_{na}(t) = \zeta(na + t) + \int_0^{na} D(na - s)dG(t + s).
\]
It is clear that the shifted functions satisfy
\[
(x_{na}(t) \wedge 1) + D_{na}(t) = \zeta_{na}(t) + \int_0^t D_{na}(t - s)dG(s).
\]
To show that \(D(\cdot)\) is non-increasing on \([na, (n + 1)a]\) is the same as to show that \(D_{na}(\cdot)\) is non-increasing on \([0, a]\). For this purpose, it is enough to verify that \(\zeta_{na}(\cdot)\) satisfy the condition (A.6). Performing integration by parts on (A.8) gives
\[
\zeta_{na}(t) = h(na + t) + (\zeta(0) - 1)^+[1 - G(na + t)] + \int_0^{na} D(na - s)dG(t + s)
\]
\[
= h(na + t) + (\zeta(0) - 1)^+[1 - G(na + t)]
\]
\[
+ D(0)G(na + t) - D(na)G(t) - \int_0^{na} G(t + s)dD(na - s).
\]
It follows from (A.1) and (A.4) that \( D(0) = (\zeta(0) - 1)^+ \), so we can write \( \zeta_{na}(\cdot) \) as
\[
\zeta_{na}(t) = h_{na}(t) + D_{na}(0)[1 - G(t)],
\]
where \( h_{na}(t) = h(na + t) + (\zeta(0) - 1)^+ - D_{na}(0) - \int_0^{na} G(t + s) dD(na - s) \). Since \( G(\cdot) \) is non-decreasing and \( D(\cdot) \) is non-increasing, the integral \( - \int_0^{na} G(t + s) dD(na - s) \) is non-increasing as a function of \( t \). So we can conclude that \( h_{na} (\cdot) \) is non-increasing, i.e. \( \zeta_{na}(\cdot) \) satisfies condition (A.6). Thus, we extend the non-increasing interval to \( [0, (n + 1)a] \). By induction, the function \( D(\cdot) \) is non-increasing on the interval \( [0, T] \).

\[\square\]

**B Glivenko-Cantelli Estimates**

An important preliminary result is the following Glivenko-Cantelli estimate. It is used in Section 5. It is convenient to state it as a general result, since the Glivenko-Cantelli estimate requires weaker conditions and gives stronger results than those in this paper.

For each \( n \), let \( \{u^n_i\}_{i \in \mathbb{Z}} \) be a sequence of i.i.d. random variables with probability measure \( \nu^F_n(\cdot) \), let \( \{v^n_i\}_{i \in \mathbb{Z}} \) be a sequence of i.i.d. random variables with probability measure \( \nu^n_G(\cdot) \). For any \( n, m \in \mathbb{Z} \) and \( l \in \mathbb{R}_+ \), define
\[
\bar{L}^F_n(m, l) = \frac{1}{n} \sum_{i=m+1}^{m+\lfloor nl \rfloor} \delta_{u^n_i}, \quad (B.1)
\]
\[
\bar{L}^G_n(m, l) = \frac{1}{n} \sum_{i=m+1}^{m+\lfloor nl \rfloor} \delta_{v^n_i}, \quad (B.2)
\]
\[
\bar{L}^n(m, l) = \frac{1}{n} \sum_{i=m+1}^{m+\lfloor nl \rfloor} \delta(u^n_i, v^n_i), \quad (B.3)
\]
where \( \delta_x \) denotes the Dirac measure of point \( x \) on \( \mathbb{R} \) and \( \delta(x, y) \) denotes the Dirac measure of point \( (x, y) \) on \( \mathbb{R} \times \mathbb{R} \). So \( \bar{L}^F_n(m, l) \) and \( \bar{L}^G_n(m, l) \) are measures on \( \mathbb{R} \) and \( \bar{L}^n(m, l) \) is a measure on \( \mathbb{R} \times \mathbb{R} \).

Denote \( C_x = (x, \infty) \), for all \( x \in \mathbb{R} \). We define two classes of testing functions by
\[
\mathcal{V} = \left\{ 1_{C_x}(\cdot) : x \in \mathbb{R} \right\},
\]
\[
\mathcal{V}_2 = \left\{ 1_{C_x \times C_y}(\cdot, \cdot) : x, y \in \mathbb{R} \right\}.
\]
It is clear that \( \mathcal{V} \) is a set of functions on \( \mathbb{R} \) and \( \mathcal{V}_2 \) is a set of functions on \( \mathbb{R} \times \mathbb{R} \). Define an envelop function for \( \mathcal{V} \) as follows. Since \( \nu^n_F \Rightarrow \nu_F \), by Skorohod representation theorem, there exists random variables \( X^n \) (with law \( \nu^n_F \)) and \( X \) (with law \( \nu_F \)), such that \( X^n \rightarrow X \) almost surely as \( r \rightarrow \infty \). Thus there exists a random variable \( X^* \) such that almost surely,
\[
X^*_r = \sup_r X^n.
\]

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Let $\nu^*_F$ be the law of $X^*$. Since $L_2(\nu^*_F)$ (the space of square integrable functions with respect to the measure $\nu^*_F$) contains continuous unbounded functions, there exists a continuous unbounded function $f_{\nu_F} : \mathbb{R}_+ \to \mathbb{R}$ that is increasing, satisfies $f_{\nu_F} \geq 1$ and $\langle f_{\nu_F}^2, \nu_F \rangle < \infty$. Similarly, based on the weak convergence $\nu^*_G \to \nu_G$, we can construct a function $f_{\nu_G}$ that is increasing, satisfies $f_{\nu_G} \geq 1$ and $\langle f_{\nu_G}^2, \nu_G \rangle < \infty$. Now, define function $\tilde{f} : \mathbb{R}_+ \to \mathbb{R}$ by $\tilde{f}(x) = \min (f_{\nu_F}(x), f_{\nu_G}(x))$ and function $\tilde{f}_2 : \mathbb{R}_+ \times \mathbb{R}_+ \to \mathbb{R}$ by $\tilde{f}_2(x,y) = \min (f_{\nu_F}(x), f_{\nu_G}(y))$ for all $x, y \in \mathbb{R}_+$. Note that we have to following properties,

$$
\tilde{f} \text{ is increasing and unbounded}, \quad (B.4)
$$

$$
\tilde{f} \leq \bar{f} \text{ for all } f \in \mathcal{V}, \quad (B.5)
$$

$$
\tilde{f} \leq \tilde{f}_2 \text{ for all } f \in \mathcal{V}_2. \quad (B.6)
$$

So we call $\tilde{f}$ and $\tilde{f}_2$ the envelop function for $\mathcal{V}$ and $\mathcal{V}_2$ respectively. Finally, let $\tilde{\mathcal{V}} = \{\tilde{f}\} \cup \mathcal{V}$ and $\tilde{\mathcal{V}}_2 = \{\tilde{f}_2\} \cup \mathcal{V}_2$.

**Lemma B.1.** Assume that

$$
\nu^*_F \to \nu_F, \quad \nu^*_G \to \nu_G \text{ as } n \to \infty.
$$

Fix constants $M, L > 0$. For all $\epsilon, \eta > 0$,

$$
\limsup_{n \to \infty} \mathbb{P}^n\left( \max_{-nM < m < nM} \sup_{l \in [0, L]} \sup_{f \in \tilde{\mathcal{V}}} \left| \langle f, \tilde{\mathcal{L}}^n_F(m, l) \rangle - l\langle f, \nu^*_F \rangle \right| > \epsilon \right) < \eta,
$$

$$
\limsup_{n \to \infty} \mathbb{P}^n\left( \max_{-nM < m < nM} \sup_{l \in [0, L]} \sup_{f \in \tilde{\mathcal{V}}} \left| \langle f, \tilde{\mathcal{L}}^n_G(m, l) \rangle - l\langle f, \nu^*_G \rangle \right| > \epsilon \right) < \eta,
$$

$$
\limsup_{n \to \infty} \mathbb{P}^n\left( \max_{-nM < m < nM} \sup_{l \in [0, L]} \sup_{f \in \tilde{\mathcal{V}}_2} \left| \langle f, \tilde{\mathcal{L}}^n(m, l) \rangle - l\langle f, \nu^*_F, \nu^*_G \rangle \right| > \epsilon \right) < \eta.
$$

This kind of results have been widely used in the study of measure valued processes, see [8, 10, 31]. The proof of the first two inequalities in the above lemma follows exactly the same way as the one for Lemma B.1 in [31], and the proof of the third inequality in the above lemma follows exactly the same as the one for Lemma 5.1 in [10]. We omit the proof for brevity. By the same reasoning as for Lemma 5.2, there exists a function $\epsilon_{GC}(\cdot)$, which vanishes at infinity such that the $\epsilon$ and $\eta$ in the above lemma can be replaced by the function $\epsilon_{GC}(n)$ for each index $n$. Based on this, we construct the following event,

$$
\Omega^n_{GC}(M, L) = \left\{ \max_{-nM < m < nM} \sup_{l \in [0, L]} \sup_{f \in \tilde{\mathcal{V}}} \left| \langle f, \tilde{\mathcal{L}}^n_F(m, l) \rangle - l\langle f, \nu^*_F \rangle \right| \leq \epsilon_{GC}(n) \right\}
$$

$$
\cap \left\{ \max_{-nM < m < nM} \sup_{l \in [0, L]} \sup_{f \in \tilde{\mathcal{V}}} \left| \langle f, \tilde{\mathcal{L}}^n_G(m, l) \rangle - l\langle f, \nu^*_G \rangle \right| \leq \epsilon_{GC}(n) \right\}
$$

$$
\cap \left\{ \max_{-nM < m < nM} \sup_{l \in [0, L]} \sup_{f \in \tilde{\mathcal{V}}_2} \left| \langle f, \tilde{\mathcal{L}}^n(m, l) \rangle - l\langle f, \nu^*_F, \nu^*_G \rangle \right| \leq \epsilon_{GC}(n) \right\}.
$$

(B.7)

It is clear that for any fixed $M, L > 0$,

$$
\lim_{n \to \infty} \mathbb{P}^n\left( \Omega^n_{GC}(M, L) \right) = 1. \quad (B.8)
$$

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Intuitively, on the event $\Omega_{GC}^n(M, L)$ (whose probability goes to 1 as $n \to \infty$ for any fixed constants $M, L$), the measures $\mathcal{L}_F^n(m, l), \mathcal{L}_G^n(m, l)$ and $\mathcal{L}_I^n(m, l)$ are very “close” to $\nu_F^n, \nu_G^n$ and $l(\nu_F^n, \nu_G^n)$, respectively.

### C An Extension of Skorohod Representation Theorem

In this section, we present a slight extension, Lemma C.1 below, of the Skorohod Representation Theorem (c.f. Theorem 3.2.2 in [26]). The proof of Lemma C.1 is built on the proof of Theorem 3.2.2 provided in the supplement of [26], with slight extension to deal with the product of two matrix spaces.

Let $(E_1, \pi_1)$ and $(E_2, \pi_2)$ be two complete and separable metric spaces. Let $(E_1 \times E_2, \pi)$ denote the product space of them, with the product metric $\pi$ obtained by the maximum metric.

**Lemma C.1.** Consider a sequence of random variables $\{(X_n, Y_n), n \geq 1\}$ in the product space $E_1 \times E_2$. If $X_n \Rightarrow X$, then there exists other random elements of $E_1 \times E_2$, $\{(\tilde{X}_n, \tilde{Y}_n), n \geq 1\}$, and $\tilde{X}$, defined on a common underlying probability space, such that

$$(\tilde{X}_n, \tilde{Y}_n) \overset{d}{=} (X_n, Y_n), n \geq 1, \quad \tilde{X} \overset{d}{=} X$$

and almost surely,

$$\tilde{X}_n \to \tilde{X} \text{ as } n \to \infty.$$

**Proof.** In order to present the proof, we first need some preliminaries. A nested family of countably partitions of a set $A$ is a collection of subsets $A_{i_1, \ldots, i_k}$ indexed by $k$-tuples of positive integers such that $\{A_i : i \geq 1\}$ is a partition of $A$ and $\{A_{i_1, \ldots, i_{k+1}} : i_{k+1} \geq 1\}$ is a partition of $A_{i_1, \ldots, i_k}$ for all $k \geq 1$ and $(i_1, \ldots, i_k) \in \mathbb{N}_+^k$. Let $P_1$ denote the probability measure on the space where $X$ lives on. Since the space $(E_1, \pi_1)$ is separable, according to Lemma 1.9 in the supplement of [26], there exists a nested family of countably partitions $\{E^1_{i_1, \ldots, i_k}\}$ of $(E_1, \pi_1)$ that satisfies

$$\text{rad}(E^1_{i_1, \ldots, i_k}) < 2^{-k}, \quad \text{rad}(E^1_{i_1, \ldots, i_k}) = 0,$$

where $\text{rad}(A)$ denotes the radius of the set $A$ in a metric space, and $\partial(A)$ denote the boundary of the set $A$. Since the space $(E_2, \pi_2)$ is separable, by the same lemma, there exists a nested sequence of countably partitions $\{E^2_{i_1, \ldots, i_k'}\}$ of $(E_2, \pi_2)$ that satisfies

$$\text{rad}(E^2_{i_1, \ldots, i_k'}) < 2^{-k'}.$$ 

Note that for space $(E_2, \pi_2)$, we only need a weaker version of Lemma 1.9 in the supplement of [26].
The first step is to use this nested sequence of countably partitions \( \{(\tilde{X}_n, \tilde{Y}_n), n \geq 1\} \) with the same distribution for each \( n \). For \( n \geq 1 \), we first construct subintervals \( I_{i_1, \ldots, i_k}^n \subseteq [0, 1) \) corresponding to the marginal probability of \( X_n \). Let \( I_1^n = [0, \mathbb{P}^n(E_1^1 \times E_2)] \) and

\[
I_i^n = \left[ \sum_{j=1}^{i-1} \mathbb{P}^n(E_{i_j}^0 \times E_2), \sum_{j=1}^{i} \mathbb{P}^n(E_{i_j}^1 \times E_2) \right], \quad i > 1,
\]

where \( \mathbb{P}^n \) is the probability measure on the space where \( (X_n, Y_n) \) lives. Let \( \{I_{i_1, \ldots, i_{k+1}}^n : i_{k+1} \geq 1\} \) be a countable partition of subintervals of \( I_{i_1, \ldots, i_k}^n \). If \( I_{i_1, \ldots, i_k}^n = [a_n, b_n) \), then

\[
I_{i_1, \ldots, i_{k+1}}^n = \left[ a_n + \sum_{j=1}^{i_k+1} \mathbb{P}^n(E_{i_1}^1 \times E_2), a_n + \sum_{j=1}^{i_k+1} \mathbb{P}^n(E_{i_1}^1 \times E_2) \right].
\]

The length of each subinterval \( I_{i_1, \ldots, i_k}^n \) is the probability \( \mathbb{P}^n(E_{i_1}^1 \times E_2) \). We then construct further subintervals \( I_{i_1, \ldots, i_k}' \subseteq I_{i_1, \ldots, i_k}^n \) corresponding to \( (X_n, Y_n) \). If \( I_{i_1, \ldots, i_k}^n = [a_n, b_n) \), then let \( I_{i_1, \ldots, i_{k+1}}^n = [a_n, a_n + \mathbb{P}^n(E_{i_1}^1 \times E_2)] \) and

\[
I_{i_1, \ldots, i_{k+1}}^n = \left[ a_n + \sum_{j=1}^{i_k+1} \mathbb{P}^n(E_{i_1}^1 \times E_2), a_n + \sum_{j=1}^{i_k+1} \mathbb{P}^n(E_{i_1}^1 \times E_2) \right], \quad i' > 1.
\]

Let \( \{I_{i_1, \ldots, i_{k+1}}^n : i_{k+1} \geq 1\} \) be countable partition of \( I_{i_1, \ldots, i_{k+1}}^n \). If \( I_{i_1, \ldots, i_{k+1}}^n = [a_n, b_n) \), then

\[
I_{i_1, \ldots, i_{k+1}}^n = \left[ a_n + \sum_{j=1}^{i_k+1} \mathbb{P}^n(E_{i_1}^1 \times E_2), a_n + \sum_{j=1}^{i_k+1} \mathbb{P}^n(E_{i_1}^1 \times E_2) \right].
\]

The length of each subinterval \( I_{i_1, \ldots, i_{k+1}}^n \) is the probability \( \mathbb{P}^n(E_{i_1}^1 \times E_2) \). Now from each nonempty subset \( E_{i_1}^1 \times E_2 \), we choose one point \((x_{i_1, \ldots, i_k}, y_{i_1, \ldots, i_k})\). For each \( n \geq 1 \) and \( k \geq 1 \), we define functions \( (x_k^n, y_k^n) : [0, 1) \to E_1 \times E_2 \) by letting \( x_k^n(w) = x_{i_1, \ldots, i_k} \) and \( y_k^n(w) = y_{i_1, \ldots, i_k} \) for \( \omega \in I_{i_1, \ldots, i_{k+1}}^n \). By the nested partition property and inequalities C.1 and C.3,

\[
\pi((x_{k}^n(\omega), x_{k+j}^n(\omega)), (x_{k}^n(\omega), x_{k+j}^n(\omega))) < 2^{-k} \quad \text{for all } j, k, n
\]

and \( \omega \in [0, 1) \). Since \((E_1 \times E_2, \pi)\) is a complete metric space, the above implies that there is \((x_n(\omega), y_n(\omega)) \in E_1 \times E_2\) such that

\[
\pi((x_{k}^n(\omega), x_{k+j}^n(\omega)), (x_{k}^n(\omega), x_{k+j}^n(\omega))) \to 0 \quad \text{as } k \to \infty.
\]

We let \( (\tilde{X}_n, \tilde{Y}_n) = (x_n, y_n) \) on \([0, 1)\) for \( n \geq 0 \).

The next step is to construct \( \hat{X} \) and show that \( \hat{X}_n \to \hat{X} \) almost surely. For each \( n \geq 1 \), let \( \mathbb{P}_1^n \) denote the marginal probability of \( X^n \). It is clear that \( I_{i_1, \ldots, i_k}^n \) is the probability \( \mathbb{P}_1^n(E_{i_1, \ldots, i_k}) \).
By (C.2), we have that \( \mathbb{P}_1^n(E_{i_1 \ldots i_k}^1) \to \mathbb{P}_1(E_{i_1 \ldots i_k}^1) \), as \( n \to \infty \). Consequently, the length of the interval \( I_{i_1 \ldots i_k}^n \) converges to the length of the interval \( I_{i_1 \ldots i_k} \), which is defined in a similar way as for \( I_{i_1 \ldots i_k}^n \) by letting

\[
I_{i_1 \ldots i_{k+1}} = \left[ a_n + \sum_{j=1}^{i_{k+1}-1} \mathbb{P}_1(E_{i_1 \ldots i_{k,j}), a_n + \sum_{j=1}^{i_{k+1}} \mathbb{P}_1(E_{i_1 \ldots i_{k,j})} \right],
\]

if \( I_{i_1 \ldots i_k} = (a_n, b_n) \). Now from each nonempty subset \( E_{i_1 \ldots i_k} \) we choose one point \( x_{i_1 \ldots i_k} \). For each \( k \geq 1 \), we define functions \( x^k : [0, 1) \to E_1 \) by letting \( x^k(\omega) = x_{i_1 \ldots i_k} \) for \( \omega \in I_{i_1 \ldots i_k}^n \). By the nested partition property and inequalities C.1,

\[
\pi_1(x^k(\omega), x^{k+j}(\omega)) < 2^{-k} \quad \text{for all } j, k
\]

and \( \omega \in [0, 1) \). Since \( (E_1, \pi_1) \) is a complete metric space, the above implies that there is \( x(\omega) \in E_1 \) such that

\[
\pi_1(x^k(\omega), x(\omega)) \to 0 \quad \text{as } k \to \infty.
\]

We let \( \tilde{X} = x \) on \([0, 1)\). Since

\[
\pi_1(\tilde{X}_n(\omega), \tilde{X}(\omega)) \leq \pi_1(\tilde{X}_n(\omega), \tilde{X}_n(\omega)) + \pi_1(\tilde{X}_n(\omega), \tilde{X}(\omega)) + \pi_1(\tilde{X}(\omega), \tilde{X}(\omega)) \\
\leq 3 \times 2^{-k},
\]

for all \( \omega \) in the interior of \( I_{i_1 \ldots i_k} \),

\[
\lim_{n \to \infty} \pi_1(\tilde{X}_n(\omega), \tilde{X}(\omega)) \leq 3 \times 2^{-k}.
\]

Since \( k \) is arbitrary, we must have \( \tilde{X}_n(\omega) \to \tilde{X}(\omega) \) as \( n \to \infty \) for all but at most countably many \( \omega \in [0, 1) \).

It remains to show that \( (\tilde{X}_n, \tilde{Y}_n) \) has the probability laws \( \mathbb{P}^n \). Let \( \tilde{\mathcal{P}} \) denote the Lebesque measure on \([0, 1)\). It suffices to show that \( \tilde{\mathcal{P}}((\tilde{X}_n, \tilde{Y}_n) \in A) = \mathbb{P}^n(A) \) for each \( A \) such that \( \mathbb{P}^n(\partial A) = 0 \). Let \( A \) be such a set. Let \( A^k \) be the union of the sets \( E_{i_1 \ldots i_k}^1 \times E_{i_1 \ldots i_k}^2 \) such that \( E_{i_1 \ldots i_k}^1 \times E_{i_1 \ldots i_k}^2 \subseteq A \) and let \( A^k \) be the union of the sets \( E_{i_1 \ldots i_k}^1 \times E_{i_1 \ldots i_k}^2 \) such that \( E_{i_1 \ldots i_k}^1 \times E_{i_1 \ldots i_k}^2 \cap A \neq \emptyset \). Then \( A^k \subseteq A \subseteq A^k \) and, by the construction above,

\[
\tilde{\mathcal{P}}((\tilde{X}_n, \tilde{Y}_n) \in A^k) = \mathbb{P}^n(A^k) \quad \text{and} \quad \tilde{\mathcal{P}}((\tilde{X}_n, \tilde{Y}_n) \in A^k) = \mathbb{P}^n(A^k)
\]

Now let \( C^k = \{ s \in E_1 \times E_2 : \pi(s, \partial A) \leq 2^{-k} \} \). Then \( A^k \setminus A^k \setminus \partial A \) as \( k \to \infty \). Since \( \mathbb{P}^n(\partial A) = 0 \) by assumption, \( \mathbb{P}^n(C^k) \downarrow 0 \) as \( k \to \infty \). Hence

\[
\tilde{\mathcal{P}}((\tilde{X}_n, \tilde{Y}_n) \in A) = \lim_{k \to \infty} \tilde{\mathcal{P}}((\tilde{X}_n, \tilde{Y}_n) \in A^k) = \lim_{k \to \infty} \mathbb{P}^n(A^k) = \mathbb{P}^n(A).
\]

Following the same way, we can show that \( \tilde{X} \) has probability law \( \mathbb{P}_1 \). \( \square \)