Abstract: This study used data from wave four of the United Kingdom (U.K.) Millennium Cohort Study to examine whether there is an individual (i.e., maternal education) and area-level social disadvantage (i.e., neighborhood deprivation) gradient to difficulties in social-emotional well-being (SEW) in 7-year-old English children. We then investigated to what extent maternal psychological distress (Kessler 6 score) explains the relationship between social disadvantage indicators and boys’ and girls’ SEW difficulties. Subjects consisted of 3661 child–mother dyads (1804 boys and 1857 girls). Results discerned gender differences in the effect social disadvantage indicators have on child SEW difficulties. Maternal education had a comparable effect on boys’ and girls’ SEW difficulties, but a steeper neighborhood deprivation gradient was evident for boys’ SEW difficulties compared to girls’ SEW difficulties. The effect of each social disadvantage indicator on boys’ and girls’ SEW difficulties was for most part direct and strong ($p \leq 0.001$) rather than through maternal psychological distress, suggesting that the theoretical framework was incomplete. Here we demonstrate that where children are positioned on the social disadvantage gradient matters greatly to their SEW. Improving the living conditions and health of mothers with psychological distress may offer a pathway to improve child SEW.

Keywords: social-emotional well-being; maternal psychological distress; neighborhood deprivation; maternal education

1. Introduction

Social-emotional well-being (SEW) relates to a child’s self-perception and their ability to interact with others. Children that experience SEW difficulties often demonstrate negative emotions and behaviors, such as anxiety, depression, and hyperactivity, and struggle to maintain friendships with peers [1]. SEW difficulties in childhood can severely impact upon children’s academic achievement and mental health in adulthood, which both convey broad negative health and economic consequences throughout the life course [2]. The prevalence of child SEW difficulties increased between 1974 and 1999 in the United Kingdom (U.K.), and in 2004, 1 in 10 children were reported to have a diagnosable SEW difficulty [3]. Evidence shows that social disadvantage is a strong predictor of child SEW difficulties [4].

Traditionally, research in this area has studied the inverse effects of individual characteristics of social disadvantage, such as low family income or maternal education, on child SEW [5,6]. However, these indicators of social disadvantage do not capture the physical neighborhood environment within which socially disadvantaged children live (i.e., high social disorder, poor housing conditions, limited access to services and aesthetic neighborhood features) which can also predispose them to SEW difficulties [7–9]. As such, each social disadvantage indicator may influence the development of child’s SEW differently. Moreover, some evidence shows gender differences in SEW difficulties [10,11]; therefore, the influence of social disadvantage on child SEW could be gender-specific. However, to
date no study has examined the influence of different indicators of social disadvantage on boys’ and girls’ SEW or determined whether a social gradient to SEW exists among boys and girls in the U.K.

The mechanisms linking social disadvantage and child SEW remain largely underexplored. Previous research in this area has predominantly been guided by the Family Investment Model (FIM) and Family Stress Model (FSM [12]). The FIM relates to the financial aspect of social disadvantage and postulates that low family income limits the purchasing of materials, experiences, and services to enhance child development and well-being [13]. The FSM rather addresses how social disadvantage influences child development and well-being through its effect on parental psychological distress and subsequent disruption of parent–child interactions/relations [14]. Previous studies investigating mechanisms underlying the association between social disadvantage and children’s SEW have been conducted mainly in the United States, which experiences greater income inequality compared to the U.K. [15,16]. Further, much of the evidence regarding SEW in U.K. children relates to younger children [5,13] who are less likely to demonstrate SEW difficulties compared to older children.

The study of child SEW is further complexed by social disadvantage and maternal psychological distress being mutually associated with child SEW, while not being independent of each other [17]. Indeed, extensive evidence shows that factors linked to social disadvantage, including low maternal income, precarious employment and housing insecurity, are strong predictors of maternal psychological distress [18]. It is important to recognize though that mothers experience particular combinations of these factors which are likely to carry varying levels of risk to their children. While the role of maternal psychological distress is considered important in the relationship between family income and early child SEW [13,19], it is unknown whether the relationship holds for older children, (boys and girls) and is consistent across different indicators of social disadvantage.

Therefore, to address these research gaps, the present study aimed to examine whether there is an individual (i.e., maternal education) and area-level social disadvantage (i.e., neighborhood deprivation) gradient to SEW in 7-year-old English children. A secondary aim was to investigate to what extent maternal psychological distress (Kessler 6 score) explains the relationship between different indicators of social disadvantage (i.e., maternal education and neighborhood deprivation) and boys’ and girls’ SEW. Firstly, it is hypothesized that children in the lowest maternal education group and those living in the most deprived neighborhoods would be at greatest risk of SEW difficulties, but the strength of these relationships would differ by gender and social disadvantage indicators. Secondly, it is hypothesized that maternal psychological distress would mediate the association between social disadvantage indicators and child SEW, and associations would vary by gender and social disadvantage indicators.

2. Materials and Methods

2.1. Participants

Data for this study were derived from wave four of the U.K. Millennium Cohort Study (MCS). The MCS is a nationally representative U.K. sample of children born between September 2000 and January 2002. The sample design allowed for over-representation of ethnic minority families and families living in areas of high neighborhood deprivation. The first survey was conducted throughout 2001–2002 when the children were aged 9 to 11 months old, and involved 18,819 children from 18,533 families. Subsequent surveys were administered at the ages of 3, 5 and 7 years. All measures were collected in the child’s home. To ensure consistency with earlier analyses [20], the present study only included English children who had complete anthropometric and physical activity data alongside the variables of interest, and for whom the main respondent to the study was the biological mother. Applying these criteria resulted in a final sample of 3661 (98.5%) child–mother dyads (1804 boys and 1857 girls). Ethical approval for the original study was granted by the Northern and Yorkshire Research Ethics Committee (07/MRE03/32).
2.2. Measures

2.2.1. Social-Emotional Well-Being Difficulties

The Strengths and Difficulties Questionnaire (SDQ) is a validated questionnaire to assess child SEW (http://www.sdqinfo.com/). The SDQ has been used as a measure of SEW difficulties globally [21]. The SDQ comprises 25 items covering five domains of behavior including hyperactivity, emotional symptoms, conduct problems, peer problems and prosocial behavior. Each item has three response categories: ‘not true’, ‘somewhat true’, or ‘certainly true’. ‘Somewhat true’ is always scored as 1, but the scoring of ‘not true’ and ‘certainly true’ responses varies across items (0 or 2). Higher scores on the prosocial behavior subscale reflect strengths, whereas higher scores on the other four subscales reflect difficulties. An overall difficulties score is derived from the sum of all scales, excluding the prosocial behavior domain [21]. Overall scores ranged from 0 to 40, with higher scores denoting more difficulties (i.e., lower SEW). SDQ difficulties were defined as a score of $\geq$14.

2.2.2. Maternal Psychological Distress

The Kessler 6 scale was used to assess maternal psychological distress [22]. Mothers reported how often over the previous 30 days they had felt everything was an effort, depressed, hopeless, restless or fidgety, worthless, and nervous. For each item, mothers indicated whether they had felt this way (1) never; (2) a little; (3) sometimes; (4) most of the time; or (5) all of the time. Responses range from 0 (none of the time) to 4 (all of the time) and are summed to produce an overall score ranging from 0–24. Serious psychological distress was defined as a score of $\geq$13.

2.2.3. Social Disadvantage

Maternal education and neighborhood deprivation were used as indicators of social disadvantage. Mothers reported their highest educational qualification achieved by the time their child was 7 years of age. Responses were classified into four categories (Higher education/degree (1); A-levels (2); O levels/GCSE A*-C (3); CSEs/GCSE D-G/none (4)). Lower levels of qualifications represented higher social disadvantage. Neighborhood deprivation was calculated from home postcodes using the 2004 English Indices of Multiple Deprivation (IMD [23]). The IMD is a U.K. Government-produced measure comprising seven areas of deprivation (income, employment, health, education, housing, environment, and crime). Decile scores were collapsed into five categories ranging from the least to most deprived neighborhoods in England. Higher decile scores represented higher social disadvantage.

2.3. Data Analysis

Analyses were conducted using SPSS v. 24 (SPSS Inc; Chicago, IL, USA) and statistical significance was set at $p \leq 0.05$. Descriptive statistics were calculated for all measured variables. Independent samples t-tests and chi-squared tests ($\chi^2$) were used to compare gender differences in continuous and categorical variables, respectively. Preliminary analyses revealed gender differences in SEW and confirmed that there was an interaction between indicators of social disadvantage (i.e., maternal education and area deprivation) and gender. Therefore, boys and girls were separated for all analyses. Gender-specific multinomial logistic regression analyses examined the likelihood of children being classified as having SEW difficulties based on maternal education and neighborhood deprivation level. The lowest level of maternal education and the highest level of neighborhood deprivation were chosen as the reference categories. For the main analyses, the outcome variable was total SDQ scores. Gender-specific mediation was assessed through regression analyses using the PROCESS macro for SPSS (http://www.processmacro.org/index.html [24]). Nonparametric bootstrapping analysis estimated direct and indirect effects in models [24]. Bootstrapping is a non-parametric resampling method that uses confidence intervals to estimate the size of indirect effects. The bootstrapping approach makes no assumption about the shape of the sampling distribution of the indirect effect and has been shown to enhance statistical power and Type I error control compared to other techniques [25].
Consistent with recommendations, 95% bias corrected, and accelerated confidence intervals were employed, and 5000 bootstrapping re-samples were run [25]. Mediation is evident when zero is not included within the lower and upper bound confidence intervals.

3. Results

Data were available for 3661 mother–child dyads (1857 girls). The descriptive characteristics of the sample and gender differences in measured variables are presented in Table 1. In this sample, 17% and 23% of children lived in the most and least deprived neighborhoods, respectively, 10% of children had a SEW difficulty, 14% of mothers had no qualification compared to 46% who had degree-level education, and 2% of mothers had psychological distress. With regards to gender differences, boys recorded higher SDQ scores than girls, and 12.3% of boys were classified as having a SEW difficulty compared to 8.5% of girls ($p \leq 0.001$).

### Table 1. Descriptive characteristics of full sample and gender differences.

| Variables                                      | All ($n = 3661$) | Boy ($n = 1804$) | Girl ($n = 1857$) |
|------------------------------------------------|------------------|------------------|-------------------|
| Age (years)                                    | 7.22 (0.25)      | 7.22 (0.25)      | 7.22 (0.25)       |
| Indices of Multiple Deprivation decile (%)     |                  |                  |                   |
| Least deprived                                 | 23.10            | 23.20            | 23.00             |
| Second                                         | 20.70            | 20.40            | 21.10             |
| Third                                          | 20.70            | 20.70            | 20.70             |
| Fourth                                         | 18.10            | 18.20            | 18.00             |
| Most deprived                                  | 17.30            | 17.50            | 17.20             |
| Maternal education level (%)                   |                  |                  |                   |
| None                                           | 26.60            | 25.30            | 27.90             |
| GCSE                                           | 14.30            | 14.70            | 13.90             |
| A-level                                        | 45.50            | 46.00            | 45.10             |
| Degree                                         | 13.50            | 14.00            | 13.10             |
| Strengths and Difficulties Questionnaire score | 6.82 (4.95)      | 7.46 (5.18)      | 6.19 (4.64) *     |
| Social-emotional well-being difficulty (%)     | 10.40            | 12.30            | 8.50 *            |
| Kessler 6 score                                | 2.84 (3.48)      | 2.86 (3.39)      | 2.81 (3.57)       |
| Serious psychological distress (%)             | 2.40             | 2.10             | 2.70              |

Social-emotional well-being difficulty $\geq 14$ on the Strengths and Difficulties Questionnaire; Serious psychological distress $\geq 13$ on the Kessler 6 scale; * $p \leq 0.001$.

3.1. Research Aim 1

Multinomial logistic regression analyses revealed that boys living in the most deprived neighborhoods were more likely to have SEW difficulties compared with boys living in the fourth to least (odds ratio, OR = 1.78; $p \leq 0.01$; Table 2), third to least (OR = 2.00; $p \leq 0.001$), second to least (OR = 2.32; $p \leq 0.001$) and least deprived neighborhoods (OR = 2.85; $p \leq 0.001$). Girls living in the most deprived neighborhoods were more likely to have SEW difficulties compared with girls living in the least deprived neighborhoods (OR = 2.57; $p \leq 0.001$).

Boys from the lowest maternal education group were more likely to have SEW difficulties compared with boys from the third to highest (OR = 2.47; $p \leq 0.001$), second to highest (OR = 2.33; $p \leq 0.001$), and highest maternal education group (OR = 3.49; $p \leq 0.001$). A steep maternal education gradient to SEW difficulties was also observed among girls. Girls from the lowest maternal education group were more likely to have SEW difficulties compared with girls from the third to highest (OR = 2.05; $p \leq 0.01$), second to highest (OR = 2.37; $p \leq 0.01$), and highest maternal education group (OR = 3.51; $p \leq 0.001$).
Table 2. Gender-specific multinomial logistic regression associations between social disadvantage indicators and social-emotional well-being difficulties (n = 3661).

| Variables                      | OR     | 95% CI    | p    |
|-------------------------------|--------|-----------|------|
| **Boys**                      |        |           |      |
| Neighborhood deprivation      |        |           |      |
| Most deprived—Reference (n = 315) | 1.78   | 1.16–2.71 | ≤0.01|
| Fourth (n = 329)              | 2.00   | 1.32–3.04 | ≤0.001|
| Third (n = 374)               | 2.32   | 1.50–3.59 | ≤0.001|
| Second (n = 367)              | 2.85   | 1.84–4.43 | ≤0.001|
| Least deprived (n = 419)      |        |           |      |
| Maternal education            |        |           |      |
| CSEs/GCSE D-G/None—Reference (n = 253) | 2.47   | 1.65–3.69 | ≤0.001|
| O-levels/GCSE A*-C (n = 456)  | 2.33   | 1.47–3.70 | ≤0.001|
| A-level (n = 265)             | 3.49   | 2.40–5.07 | ≤0.001|
| Degree (n = 830)              |        |           |      |
| **Girls**                     |        |           |      |
| Neighborhood deprivation      |        |           |      |
| Most deprived—Reference (n = 319) | 1.02   | 0.62–1.68 | 0.94 |
| Fourth (n = 334)              | 0.97   | 0.60–1.57 | 0.91 |
| Third (n = 384)               | 1.55   | 0.92–2.62 | 0.10 |
| Second (n = 392)              | 2.57   | 1.44–4.39 | ≤0.001|
| Least deprived (n = 428)      |        |           |      |
| Maternal education            |        |           |      |
| CSEs/GCSE D-G/None—Reference (n = 243) | 2.05   | 1.31–3.20 | ≤0.01|
| O-levels/GCSE A*-C (n = 518)  | 2.37   | 1.36–4.13 | ≤0.01|
| A-level (n = 259)             | 3.51   | 2.25–5.48 | ≤0.001|
| Degree (n = 837)              |        |           |      |

CI, confidence interval; OR, odds ratio; Social-emotional well-being difficulties = score of ≥14 on the Strengths and Difficulties Questionnaire.

3.2. Research Aim 2

Figure 1 displays unstandardized regression coefficients for the gender-specific mediation models. In both models, neighborhood deprivation and maternal education were the independent variables and total SDQ scores the dependent variable. Maternal education was the mediator (panel A and B). Results of the indirect effects for boys and girls are presented in Table 3. Figure 1 shows that neighborhood deprivation and maternal education were related to maternal psychological distress for boys ($\beta = 0.28$, $p \leq 0.0001$; $\beta = -0.33$, $p \leq 0.0001$) and girls ($\beta = 0.24$, $p \leq 0.0001$; $\beta = -0.27$, $p \leq 0.0003$). Maternal psychological distress was related to boys’ ($\beta = 0.55$, $p \leq 0.001$), and girls’ SDQ scores ($\beta = 0.38$, $p \leq 0.0001$).

In the neighborhood deprivation model for boys (Figure 1, panel A), the bootstrap procedure revealed that the total effect of neighborhood deprivation on SDQ scores was significant ($\beta = 0.58$, $p \leq 0.0001$). When the mediator was included in the model, the direct effect of neighborhood deprivation on boys’ SDQ scores reduced but remained significant, suggesting partial mediation ($\beta = 0.42$, $p \leq 0.0001$). Maternal psychological distress displayed an indirect effect, $\beta = 0.15$; 95% confidence interval (CI) = 0.09–0.22; $p \leq 0.0001$ (Table 3). Therefore, the effect of neighborhood deprivation on boys’ SDQ scores was partially mediated by maternal psychological distress. In the maternal education model for boys (Figure 1, panel B), the bootstrap procedure revealed that the total effect of maternal education on SDQ scores was significant ($\beta = -0.94$, $p \leq 0.0001$). When the mediator was included in the model, the direct effect of maternal education on boys’ SDQ scores reduced but remained significant, suggesting partial mediation ($\beta = -0.76$, $p \leq 0.0001$). Maternal psychological distress displayed an indirect effect, $\beta = -0.18$; 95% CI = -0.27–−0.10; $p \leq 0.0001$ (Table 3). Therefore, the effect of maternal education on boys’ SDQ scores was partially mediated by maternal psychological distress.
Figure 1. Cont.
Figure 1. (a,b) Regression models predicting social-emotional well-being difficulties for boys. Values signify unstandardized regression coefficients. The direct effect of neighborhood deprivation and maternal education on social-emotional well-being are outside parentheses. The total effect is inside parentheses. *** p ≤ 0.001; (c,d) Regression models predicting social-emotional well-being difficulties for girls. Values signify unstandardized regression coefficients. The direct effect of neighborhood deprivation and maternal education on social-emotional well-being are outside parentheses. The total effect is inside parentheses. *** p ≤ 0.001.
Table 3. Indirect effects of neighborhood deprivation and maternal education on child social-emotional well-being difficulties through maternal psychological distress by gender ($n = 3661$).

|               | Boys                                                                 | Girls                                                                    |
|---------------|----------------------------------------------------------------------|-------------------------------------------------------------------------|
| **Neighbourhood deprivation—Social-emotional well-being difficulties** | **Normal Theory Tests** | **Normal Theory Tests** |
| Total effect  | Bootstrap Effect 0.58, 95% CI 0.41–0.74 | Bootstrap Effect 0.55, 95% CI 0.40–0.69 |  
| Maternal psychological distress | 0.15, 95% CI 0.09–0.22 | 0.09, 95% CI 0.05–0.14 | 
| **Maternal education—Social-emotional well-being difficulties** | **Normal Effect 0.15, 95% CI 0.09–0.22, SE 0.03, z 4.72, p <0.0001** | **Normal Effect −0.18, 95% CI −0.27–−0.10, SE −0.18, z −4.48, p <0.0001** |
| Total effect  | −0.94, 95% CI −1.15–−0.73 | −0.80, 95% CI −0.99–−0.62 |  
| Maternal psychological distress | −0.18, 95% CI −0.27–−0.10 | −0.10, 95% CI −0.17–−0.04 | 

Note: Bootstrap-generated confidence intervals. CI = confidence interval; SE, standard error.
In the neighborhood deprivation model for girls (Figure 1, panel C), the bootstrap procedure revealed that the total effect of neighborhood deprivation on SDQ scores was significant ($\beta = 0.55$, $p \leq 0.0001$). When the mediator was included in the model, the direct effect of neighborhood deprivation on girls’ SDQ scores reduced but remained significant, suggesting partial mediation ($\beta = 0.45$, $p \leq 0.0001$). Maternal psychological distress displayed an indirect effect, $\beta = 0.09$; 95% CI = 0.05–0.14; $p \leq 0.0001$ (Table 3). Therefore, the effect of neighborhood deprivation on girls’ SDQ scores was partially mediated by maternal psychological distress. In the maternal education model for girls (Figure 1, panel D), the bootstrap procedure revealed that the total effect of maternal education on SDQ scores was significant ($\beta = -0.80$, $p \leq 0.0001$). When the mediator was included in the model, the direct effect of maternal education on girls’ SDQ scores reduced but remained significant, suggesting partial mediation ($\beta = -0.70$, $p \leq 0.0001$). Maternal psychological distress displayed an indirect effect, $\beta = -0.10$; 95% CI = -0.17–-0.04; $p = 0.0004$ (Table 3). Therefore, the effect of maternal education on girls’ SDQ scores was partially mediated by maternal psychological distress.

4. Discussion

This study is the first to examine associations between individual and area-level indicators of social disadvantage and child SEW difficulties and investigate the mediating effect of maternal psychological distress on the social disadvantage–child SEW difficulties relationship. In the present study, both indicators of social disadvantage were strongly positively associated with child SEW difficulties, and further analyses revealed marked child SEW inequalities between the most and least socially disadvantaged categories. This finding supplements previous work which evidenced a steep social gradient to child obesity in England based on the same social disadvantage indicators used in the present study [20]. Together the findings underscore the importance of assessing child health across a range of social disadvantage indices. Failure to do so could potentially mask the pervasive effect of social disadvantage on child health. We also revealed that the strength of association between social disadvantage and child SEW difficulties varied by social disadvantage indicator. Indeed, aside from maternal education being a stronger predictor of boys’ and girls’ SEW difficulties compared to neighborhood deprivation, a steeper gradient to child SEW difficulties was also evident for maternal education relative to neighborhood deprivation. This finding could be due to maternal education being a stronger indicator of low income and financial hardship, which are central facets of the FIM [12]. Low maternal education has been found to be a strong predictor of child SEW difficulties in previous U.K. [26] and European research [27]. However, here we draw specific attention to the social gradient and child SEW. Where children are positioned on the social disadvantage gradient matters greatly to their SEW.

A second study objective was to examine the mediating effect of maternal psychological distress on the relationship different indicators of social disadvantage and child SEW difficulties. Our hypothesis was that the increased risk of child SEW difficulties may be explained by maternal psychological distress. This hypothesis was partially supported and corroborates somewhat with the FSM [12]. However, the effect of each social disadvantage indicator on child SEW difficulties was for most part direct and strong rather than through maternal psychological distress, suggesting that the theoretical framework was incomplete. While no previous studies have examined the exact mediation mechanisms that were the focus of the present study, the results from our study are consistent with younger samples. For example, previous U.K. research involving young children revealed that the social disadvantage–child SEW difficulties relationship remained significant after accounting for parental stress and other factors known to influence child behavior [19]. Together, these findings indicate that additional research is needed to explain the processes by which different indicators of social disadvantage affect child SEW difficulties. Future research should investigate the mediating effects of child stress and lifestyle factors (e.g., dietary intake) and/or neighborhood attributes on the social disadvantage–child SEW difficulties relationship.
A novel element of this study was the examination of gender-specific relationships between social disadvantage, maternal psychological distress and child SEW difficulties. The gender-specific multinomial logistic regression analyses discerned gender differences in the effect social disadvantage indicators have on child SEW difficulties. Interestingly, while maternal education had a comparable effect on boys’ and girls’ SEW difficulties, a steeper neighborhood deprivation gradient was evident for boys’ SEW difficulties compared to girls’ SEW difficulties. This finding would suggest that position on the deprivation gradient matters more for boys’ SEW compared to girls’ SEW. Moreover, in both mediation models, larger total and indirect effects were observed for boys compared to girls. Together, these findings propose that the independent and combined effect of social disadvantage and maternal psychological distress predispose boys to SEW difficulties more so than girls. Presently, it is largely unknown whether differential effects exist for neighborhood and child SEW relationships. Quasi-experimental studies such as the Moving to Opportunity program gave residents living in some of America’s most socially disadvantaged neighborhoods the opportunity to move to more affluent neighborhoods [28]. The program resulted in reduced rates of conduct disorder among girls but increased rates of depression and conduct disorder among boys. These findings suggest that gendered intervention programs are needed and that boys may need additional program support than girls.

It is difficult to fully compare our findings to earlier work as these studies involved younger samples and did not explore gender differences [5,13,19]. However, a review study reported that maternal depression was more strongly associated with internalizing problems (e.g., emotional and peer symptoms) in girls than boys [29]. Here we assessed the influence of maternal psychological distress on total SDQ scores rather than specific internalizing and externalizing problems which may have contributed to the heterogeneity in results. There are a multitude of adversities associated with social disadvantage that collectively contribute to its pervasive impact on child SEW, and specifically, internalizing and externalizing problems [7,30]. A majority of interventions to improve child SEW have targeted only one social disadvantage risk factor (e.g., family income, neighborhood) and efforts targeting a range of social disadvantage risk factors are likely to be needed to bring about greater and more continued change [16]. However, additional research is needed to understand how the various individual, family, and neighborhood risk factors interact to create unconducive environments for boys’ and girls’ development and well-being.

Consistent with prior research this study revealed that maternal psychological distress was strongly influenced by social disadvantage [31], especially maternal education level, and strongly predicted both boys’ and girls’ SEW difficulties. Education level influences a mother’s employment opportunities which in part determines their income level. Although all mothers are likely to experience some level of stress, mothers with less education, income, and social status have fewer resources to mitigate that stress in comparison to mothers with higher education, income, and social status. Maternal psychological distress makes optimal parenting more challenging and has been shown to impact negatively on child well-being [17,29]. Compared to low education mothers, mothers with a high level of education are more likely to be cognizant of normative child development milestones as well as approaches appropriate to fulfilling the social-emotional and cognitive needs of their children as they develop [32]. Moreover, it is likely that highly educated mothers are more effective in providing the social support required to cope with the negative effect maternal psychological distress has on their child’s SEW [33]. Mother’s educational level may also directly influence father parenting behavior and father-child relations, given that a mother’s educational level influences their attribution of children’s behavior which in turn affects the quality of their parenting decisions and behavior [32,34]. Children of low-educated mothers experiencing maternal psychological distress are at the greatest risk of SEW difficulties and thus represent an important target group for future child SEW interventions.

Overall, the present study demonstrates the need for a broader recognition of maternal mental health as a foundation for children’s well-being. In this study, mothers with the lowest education and those living in the most deprived neighborhoods were most likely to report psychological distress and child SEW difficulties. Together these results suggest that the poorer well-being of mothers
experiencing psychological distress and their children can potentially be attributed to their poorer living conditions, rather than mothers’ well-being per se. The results are consistent with a socio-ecological approach to child- and maternal health promotion [35], whereby changes to context can improve multiple family health outcomes [36]. Traditionally, U.K. policy efforts to reduce social inequalities in child health have dedicated resources on children themselves. The results of this study suggest that a more family-centered approach to child health promotion whereby efforts are focused on enhancing mothers’ mental health may benefit children’s SEW.

There are some study limitations to acknowledge. Firstly, maternal reported SDQ scores may have been subject to measurement error and social desirability bias and may reflect gendered stereotypes of children’s behavior. Furthermore, we did not include all key pathways of the FSM (i.e., parenting practices) which may also influence the social disadvantage–child SEW relationship. Moreover, the study design was cross-sectional and does not determine causality. It is also important to note that the results of this study are based on data collected in 2007–2008 and may underestimate the current magnitude of health inequalities among English children and their mothers. For example, since 2010 a range of Conservative Government-led austerity measures (including cuts to children and family support services, welfare benefits, and tax credits) have been implemented, which have in most part hit the poorest areas of England hardest [37]. Coupled with these policy changes, the U.K. has experienced a rapid rise in precarious employment [38] and housing insecurity [39], all of which are strong risk factors for adult psychological distress [40,41]. It is not unreasonable to assume that these societal changes will have impacted most on the living conditions and consequent health of children and families living in the most deprived areas of England, and in turn widened health inequalities.

5. Conclusions

Our study evidences marked social disadvantage inequalities in English children’s SEW at age 7 years. The results of this study reinforce strong associations between individual and area-level indicators of social disadvantage and child SEW difficulties. They show that where children are positioned on the social disadvantage gradient matters greatly to their SEW. We also demonstrate the influence maternal psychological distress has on the social disadvantage–child SEW difficulties relationship. Improving the living conditions and health of mothers with psychological distress may offer a pathway to improve child SEW.

Author Contributions: R.J.N. designed the study, analyzed and interpreted the data and drafted the manuscript. S.J.F. critically revised the manuscript. Both authors read and approved the final manuscript.

Funding: This research received no external funding and was supported by Edge Hill University.

Acknowledgments: The authors are grateful to the participating families of the U.K. Millennium Cohort Study, Centre for Longitudinal Studies, UCL Institute of Education for the use of these data and to the U.K. Data Archive and U.K. Data Service for making them available. However, they bear no responsibility for the analysis or interpretation of these data.

Conflicts of Interest: The authors declare no conflict of interest.

References
1. National Institute for Health and Care Excellence. Social and Emotional Wellbeing for Children and Young People; National Institute for Health and Care Excellence: London, UK, 2017.
2. Marmot, M. Fair Society, Healthy Lives: The Marmot Review: Strategic Review of Health Inequalities in England Post 2010; The Marmot Review: London, UK, 2010.
3. Murphy, M.; Fonagy, P. Chief Medical Officer annual report 2012: children and young people’s health; Department of Health: London, UK, 2013; Chapter 10.
4. Reiss, F. Socioeconomic inequalities and mental health problems in children and adolescents: A systematic review. Soc. Sci. Med. 2013, 90, 24–31. [CrossRef] [PubMed]
5. Fitzsimons, E.; Goodman, A.; Kelly, E.; Smith, J.P. Poverty dynamics and parental mental health: Determinants of childhood mental health in the UK. Soc. Sci. Med. 2017, 175, 43–51. [CrossRef] [PubMed]
6. Rijlaarsdam, J.; Stevens, G.W.J.M.; van der Ende, J.; Hofman, A.; Jaddoe, V.W.; Mackenbach, J.P.; Verhulst, F.C.; Tiemeyer, H. Economic Disadvantage and Young Children’s Emotional and Behavioral Problems: Mechanisms of Risk. *J. Abnorm. Child Psychol.* 2013, 41, 125–137. [CrossRef] [PubMed]

7. Minh, A.; Muhajarine, N.; Janus, M.; Brownell, M.; Guhn, M. A review of neighborhood effects and early child development: How, where, and for whom, do neighborhoods matter? *Health Place* 2017, 46, 155–174. [CrossRef] [PubMed]

8. Noonan, R.J.; Boddy, L.M.; Knowles, Z.R.; Fairclough, S.J. Cross-sectional associations between high-deprivation home and neighbourhood environments, and health-related variables among Liverpool children. *BMJ Open* 2016, 6, e008693. [CrossRef] [PubMed]

9. Weitzman, M.; Baten, A.; Rosenthal, D.G.; Hoshino, R.; Tohn, E.; Jacobs, D.E. Housing and Child Health. *Curr. Probl. Pediatr. Adolesc. Health Care* 2013, 43, 187–224. [CrossRef] [PubMed]

10. Psychogiou, L.; Moberly, N.J.; Parry, E.; Nath, S.; Kallitsoglou, A.; Russell, G. Parental depressive symptoms, children’s emotional and behavioural problems, and parents’ expressed emotion—Critical and positive comments. *PLoS ONE* 2017, 12, e0183546. [CrossRef] [PubMed]

11. Zahn-Waxler, C.; Shirtcliff, E.A.; Marceau, K. Disorders of childhood and adolescence: Gender and psychopathology. *Annu. Rev. Clin. Psychol.* 2008, 4, 275–303. [CrossRef] [PubMed]

12. Conger, R.; Elder, G.H. *Families in Troubled Times: Adapting to Change in Rural America*; Aldine de Gruyter: New York, NY, USA, 1994.

13. Kiernan, K.E.; Huerta, M.C. Economic deprivation, maternal depression, parenting and children’s cognitive and emotional development in early childhood. *Br. J. Sociol.* 2008, 59, 783–806. [CrossRef] [PubMed]

14. Elgar, F.J.; Mills, R.S.L.; McGrath, P.J.; Waschbusch, D.A.; Brownridge, D.A. Maternal and paternal depressive symptoms and child maladjustment: The mediating role of parental behavior. *J. Abnorm. Child Psychol.* 2007, 35, 943–955. [CrossRef] [PubMed]

15. Organisation for Economic Co-Operation and Development. *An Overview of Growing Income Inequalities in OECD Countries: Main Findings*; Organisation for Economic Co-Operation and Development: Paris, France, 2011.

16. Yoshikawa, H.; Aber, J.L.; Beardslee, W.R. The Effects of Poverty on the Mental, Emotional, and Behavioral Health of Children and Youth: Implications for Prevention. *Am. Psychol.* 2012, 67, 272–284. [CrossRef] [PubMed]

17. Hindmarsh, G.; Llewellyn, G.; Emerson, E. The Social-Emotional Well-Being of Children of Mothers with Intellectual Impairment: A Population-Based Analysis. *J. Appl. Res. Intellect. Disabil.* 2017, 30, 469–481. [CrossRef] [PubMed]

18. Wickham, S.; Whitehead, M.; Taylor-Robinson, D.; Barr, B. The effect of a transition into poverty on child and maternal mental health: A longitudinal analysis of the UK Millennium Cohort Study. *Lancet Public Health* 2017, 2, e141–e148. [CrossRef]

19. Violato, M.; Petrou, S.; Gray, R.; Redshaw, M. Family income and child cognitive and behavioural development in the United Kingdom: Does money matter? *Health Econ.* 2011, 20, 1201–1225. [CrossRef] [PubMed]

20. Noonan, R.J.; Fairclough, S.J. Is there a deprivation and maternal education gradient to child obesity and moderate-to-vigorous physical activity in England? Findings from the Millennium Cohort Study. *Pediatr. Obes.* 2018, 13, 458–464. [CrossRef] [PubMed]

21. Goodman, A.; Goodman, R. Strengths and difficulties questionnaire as a dimensional measure of child mental health. *J. Am. Acad. Child Adolesc. Psychiatry* 2009, 48, 400–403. [CrossRef] [PubMed]

22. Kessler, R.C.; Andrews, G.; Colpe, L.J.; Hiripi, E.; Mroczek, D.K.; Normand, S.L.; Walters, E.E.; Zaslavsky, A.M. Short screening scales to monitor population prevalences and trends in non-specific psychological distress. *Psychol. Med.* 2002, 32, 959–976. [CrossRef] [PubMed]

23. Department for Communities and Local Government. *Indices of Deprivation 2004*; Department for Communities and Local Government: London, UK, 2005.

24. Hayes, A.F. *Introduction to Mediation, Moderation, and Conditional Process Analysis: A Regression-Based Approach*; Guilford Press: New York, NY, USA, 2013.

25. Hayes, A.F. Beyond Baron and Kenny: Statistical mediation analysis in the new millennium. *Commun. Monogr.* 2009, 76, 408–420. [CrossRef]

26. Marryat, L.; Thompson, L.; Minnis, H.; Wilson, P. Exploring the social, emotional and behavioural development of preschool children: Is Glasgow different? *Int. J. Equity Health* 2015, 14, 3. [CrossRef] [PubMed]
27. Barriuso-Lapresa, L.; Hernando-Arizatea, L.; Rajmil, L. Social inequalities in mental health and health-related quality of life in children in Spain. *Pediatrics* **2012**, *130*, 528–535. [CrossRef] [PubMed]

28. Osypuk, T.L.; Tchetgen, E.J.; Acevedo-Garcia, D.; Earls, F.J.; Lincoln, A.; Schmidt, N.M.; Glymour, M.M. Differential mental health effects of neighborhood relocation among youth in vulnerable families: Results from a randomized trial. *Arch. Gen. Psychiatry* **2012**, *69*, 1284–1294. [CrossRef] [PubMed]

29. Goodman, S.H.; Rouse, M.H.; Connell, A.M.; Broth, M.R.; Hall, C.M.; Heyward, D. Maternal Depression and Child Psychopathology: A Meta-Analytic Review. *Clin. Child Fam. Psychol. Rev.* **2011**, *14*, 1–27. [CrossRef] [PubMed]

30. Hedman, L.; Manley, D.; van Ham, M.; Östh, J. Cumulative exposure to disadvantage and the intergenerational transmission of neighbourhood effects. *J. Econ. Geogr.* **2015**, *15*, 195–215. [CrossRef]

31. Mensah, F.K.; Kiernan, K.E. Maternal general health and children’s cognitive development and behaviour in the early years: Findings from the Millennium Cohort Study. *Child Care Health Dev.* **2010**, *37*, 44–54. [CrossRef] [PubMed]

32. Bornstein, M.H.; Cote, L.R.; Haynes, O.M.; Hahn, C.S.; Park, Y. Parenting knowledge: Experiential and sociodemographic factors in European American mothers of young children. *Dev. Psychol.* **2010**, *46*, 1677–1693. [CrossRef] [PubMed]

33. Feinstein, L.; Duckworth, K.; Sabates, R. *Education and the Family: Passing Success Across the Generations*; Routledge: London, UK, 2008.

34. Park, J.L.; Johnston, C.; Colalillo, S.; Williamson, D. Parents’ Attributions for Negative and Positive Child Behavior in Relation to Parenting and Child Problems. *J. Clin. Child Adolesc. Psychol.* **2016**, *12*, 1–13. [CrossRef] [PubMed]

35. Stokols, D. Establishing and maintaining healthy environments: Toward a social ecology of health promotion. *Am. Psychol.* **1992**, *47*, 6–22. [CrossRef] [PubMed]

36. Gross, H.E.; Shaw, D.S.; Moilanen, K.L. Reciprocal associations between boys’ externalizing problems and mothers’ depressive symptoms. *J. Abnorm. Psychol.* **2008**, *36*, 693–709. [CrossRef] [PubMed]

37. UK Parliament. Impact of the Spending Review on Health and Social Care [Internet]. 2016. Available online: https://www.publications.parliament.uk/pa/cm201617/cmselect/cmhealth/139/13902.htm?utm_source=139&utm_medium=fullbullet&utm_campaign=Modulereports (accessed on 10 November 2018).

38. TUC. The Millions Trapped in Insecure Work Deserve a New Deal [Internet]. 2018. Available online: https://www.tuc.org.uk/blogs/millions-trapped-insecure-work-deserve-new-deal (accessed on 10 November 2018).

39. Shelter. *Unsettled and Insecure: The Toll Insecure Private Renting is Taking on English Families*; Shelter: London, UK, 2017.

40. Rajani, N.B.; Giannakopoulos, G.; Filippidis, F.T. Job insecurity, financial difficulties and mental health in Europe. *Occup. Med.* **2016**, *66*, 681–683. [CrossRef] [PubMed]

41. Suglia, S.F.; Duarte, C.S.; Sandel, M.T. Housing Quality, Housing Instability, and Maternal Mental Health. *J. Urban Health* **2011**, *88*, 1105–1116. [CrossRef] [PubMed]

© 2018 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (http://creativecommons.org/licenses/by/4.0/).