Negative Life Events Vary by Neighborhood and Mediate the Relation between Neighborhood Context and Psychological Well-Being

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Abstract

Researchers have speculated that negative life events are more common in troubled neighborhoods, amplifying adverse effects on health. Using a clustered representative sample of Chicago residents (2001–03; n = 3,105) from the Chicago Community Adult Health Survey, we provide the first documentation that negative life events are highly geographically clustered compared to health outcomes. Associations between neighborhood context and negative life events were also found to vary by event type. We then demonstrate the power of a contextualized approach by testing path models in which life events mediate the relation between neighborhood characteristics and health outcomes, including self-rated health, anxiety, and depression. The indirect paths between neighborhood conditions and health through negative life event exposure are highly significant and large compared to the direct paths from neighborhood conditions to health. Our results indicate that neighborhood conditions can have acute as well as chronic effects on health, and that negative life events are a powerful mechanism by which context may influence health.

Introduction

Negative life events are powerful potential triggers in producing stress-related health effects [1,2]. These stressful and potentially traumatic events occur more frequently and earlier among those with low socioeconomic status [2–7]. In a review of mechanisms linking psychological factors and health, Miller and colleagues [8] hypothesized that negative life events may also be more common in disadvantaged neighborhoods. Variation in neighborhood conditions due to segregation may also be a key contributor to health disparities [9,10]. Context may to some extent itself predict or condition stress, emotions, and coping behaviors directly [11–15]. We provide the first evidence that neighborhood conditions are related to the prevalence of recent negative life events, and demonstrate the importance of this link for psychological well-being.

Most survey-based research on life events has implicitly treated life event risk as a risk faced by individuals. We argue that individuals are embedded in communities, that some communities experience disproportionate social, demographic, and physical hazards (6–8), and that these hazards are intertwined. The phrase “riskscape” [14] conveys the idea that particular physical and social risks (or resources) in a place are not independent but rather interact and accumulate. This growing appreciation that contextual risk is more than the sum of risks from individual toxicants or stressors has given rise to a new emphasis on documenting the nature of these cumulative exposures [15]. Just as a concentration of poverty, social dysfunction, and crime in a community may lead to worse health outcomes [16], we expect that certain neighborhood conditions may independently increase the risk of future negative life events. Meanwhile, negative life events may also have triggered downward social mobility. Downward social mobility may co-occur with moves to worse neighborhoods or staying in declining neighborhoods. Hardships and traumatic experiences in certain communities may be associated with subsequent social, psychological, and health problems. Thus, we hypothesize strong links between recent negative life events and neighborhood conditions.

Life events are not purely exogenous shocks but rather are correlated with exposures to other risks and resources in one’s environment. Variation in recent experience of negative life events is one important mechanism by which deleterious neighborhood effects may relate to downstream health outcomes. However, existing research has not yet assessed whether neighborhood conditions are linked with negative life event experience. Using data from a representative sample of adults clustered in 343 contiguous Chicago neighborhoods, we report the prevalence of a variety of negative life events in a representative urban sample and...
Life Events in Spatial Context

Importance of Life Events. Exposure to negative life events has been associated with a variety of adverse physical and psychological health outcomes, including symptoms of depression [17] and post-traumatic stress disorder (PTSD) [18], reduced life satisfaction [19], inflammation [20], and chronic physical conditions [21]. In addition, research has shown that traumatic experiences over the life course can exert greater adverse effects on psychological and physical health outcomes compared to single incident events [22–24]. Given that the majority of individuals exposed to stressful and traumatic events experience more than one event during their lifetime [25], it is important to understand the cumulative effects of exposure to multiple negative life events as well as the differential impact of exposure to various types of events separately.

Life Events in Context. Life events have typically been viewed as individual risks. However, life event risks and experiences take place in physical locations which carry different risks for exposure to certain types of events. Negative life events are more common in places that are undergoing periods of violence and hardship, such as war and natural disaster. Dropping out of school and teenage childbearing are also much more common in troubled neighborhoods [26]. Durkheim’s classic work on suicide [27] demonstrated differences in suicide rates between Catholic and Protestant places. Also, characteristics of individuals’ social networks (embedded in place) may increase the risk of experiencing a negative life event. For instance, older adults are particularly likely to experience the death of social contacts, simply because many of the people they know are also older adults.

Thus, life events occur within socially- and spatially-variant contexts. While a few papers have discussed a specific context in which life event data were collected (place; i.e. [28]), research has rarely considered how various conditions in locations in which respondents spend time (space) may relate to their experience of life events. Life experiences may vary across neighborhoods as a result of variations in worldviews, economic opportunities, physical safety, social support, and access to resources (e.g., transportation, health care, child care). Certain aspects of family dynamics, such as the role of older adults in the family structure and the amount of responsibility placed on children and young adults, may also vary culturally across neighborhood contexts. These differences may in turn influence the risk (and impact) of negative life events. Only one prior study looked at life events in context, finding neighborhood disadvantage and disorder amplified the association between negative life events and the onset of major depression among African-American women [29].

Health in Context. Advancing research concerning contextual effects on health will require greater attention to concrete mechanisms by which specific features of neighborhoods “get under the skin” [9,30]. One way to begin looking for contextual causal mechanisms is to investigate the geographic clustering of risks, resources, and outcomes. When residential neighbors have more similar values of a variable than those who live farther away (geographic clustering) [31], this suggests the potential for a neighborhood mechanism. To some extent, geographic clustering can be spurious – the result of composition effects. For instance, neighborhoods populated by older adults will have higher mortality because older adults have higher mortality. Neighborhood effects could also be due to selective migration related to the variable being studied. For example, the life event “death of spouse” would be more common in Sunbelt retirement towns because older adults migrate there in later life expecting to live well, not because retirement towns are dangerous. Negative life events may also be contagious, rather than causal or compositional, such as when gang violence results in a string of reprisals. Once geographic clustering has been documented, identifying which of these mechanisms is at play and to what extent should be an important research goal.

To illustrate, crime rates vary substantially across intra-urban neighborhoods [32], and living in a high-crime area may influence health in multiple ways. The effort to avoid crime may influence health behaviors – to avoid being victimized, residents may avoid walking for exercise or transportation and spending time outdoors. Continual fear of crime may also result in chronic activation of stress processes, which may lead to physiological dysregulation and later disease [33]. Fear of crime may also influence social behaviors such as avoiding interacting with neighbors, community participation, and moving away, triggering a reciprocal effect such that fear of crime induces conditions which may foster social disorganization [34], which in turn leads to higher crime rates. Although prior research has focused on chronic pathways like fear, our goal is to examine the acute effects of neighborhood conditions, which are experienced only in the instance in which the feared event occurs (e.g. a resident is a victim of a crime) (Figure 1). This distinction between chronic and acute effects applies to many aspects of place, such as fear of police vs. being unnecessarily frisked, traffic stress vs. having an accident, or anxiety while home alone vs. being robbed. It is likely that the health effects of actually experiencing the feared event (the acute exposure) is much greater than the effect of the fear (the chronic exposure), even though far more people are exposed to the chronic than the acute stressor.

Pathways. Figure 2 illustrates potential pathways between neighborhood characteristics and health outcomes through life events. Crucially, we do not know the comparative amount that stress- or event-based pathways influence downstream health. As a result, we might overestimate the contextual effect of fear of
negative events because sometimes those events actually occur, and both fear of and exposure to the event may influence the same health outcome. Thus life events might mediate and or moderate neighborhood conditions. For example, a higher crime rate increases the risk of victimization, and having been victimized might increase the stress response to exposure to further crime risk. The pathway through life events may also suppress a direct effect. Failing to account for acute as well as chronic effects might thus result in model misspecification. Most neighborhood studies include individual controls and neighborhood predictors, and many recognize potential confounding between the two levels (i.e., between individual and neighborhood income/wealth and race/racial composition [33]). But little attention has been given to potential confounding between individual and contextual stressors. Researchers interested in stress mechanisms typically introduce a neighborhood measure into a model, understanding the direct pathway to be the total effect. Our main goal is to draw attention to a potential additional pathway not previously emphasized in the literature. Increased risk of a negative life event having occurred to respondents in a certain neighborhood (causal or spurious) may result in an individual experiencing the event, and experience of the event may influence their health in an acute way that is distinct from the chronic effect. Alternatively, a negative life event may propel an individual into worse contexts. In other words, life events may mediate/confound/suppress the relationship between neighborhood characteristics and health outcomes. (Confounding and mediation are computationally equivalent in cross-sectional models, and attribution of the effect type depends on the researchers' understanding of the pathways involved [36].)

Specifically, our aim is to illustrate that when neighborhood predictors are entered into a model predicting a health outcome, those neighborhood predictors are correlated with the recent experiences of individuals in prior contexts. For this reason we use retrospective rather than prospective data, and focus on where respondents live at the time of data collection rather than when the event occurred, consistent with the data framework used in research on stress and neighborhood effects on health.

Materials and Methods

Sample

To assess the association of negative life events with neighborhood characteristics and health, we use a multilevel probability sample of 3,105 adults age 18 and older, with analyses weighted to represent the Chicago population in 2000. The Chicago Community Adult Health Study (CCHAS) consists of face-to-face interviews (71.8% response rate), systematic social observation, a community survey, and linkage with archival data. The survey elicits individual level data about socioeconomic, psychosocial, behavioral factors, health, and perceived social and physical characteristics of neighborhoods. Neighborhoods in the CCHAS are operationalized as clusters (NGs) of contiguous census tracts, based on the clustered sampling framework of the Project on Human Development in Chicago Neighborhoods [37] and reflecting physical barriers, local cultural knowledge, and cluster analyses of census data so that the NGs are relatively homogeneous and cover the entire city [38]. On average 9 respondents live in each of the 343 NGs. Analyses are weighted to represent Chicago’s 2000 Census population in terms of age, race/ethnicity, and sex.

The sociodemographic composition of the sample has been reported elsewhere [13]. There are substantial numbers of minorities and a broad range of adult ages and socioeconomic statuses [39]. Data collection for the CCHAS was approved under the University of Michigan Health Sciences and Behavioral Sciences Institutional Review Boards. Respondents gave written informed consent, and data were deidentified prior to public release.

Measurement

Negative Life Events. Respondents were asked about their experience of 16 specific negative life events in the last 5 years (Table 1). CCHAS selected events using a preliminary field study based on the Duke Life Events Inventory [1,40]. To examine the extent to which neighborhood conditions differentially relate to various types of life events, events were classified into several dichotomous categories (Table 1). First, events were classified as respondent-directed or other-directed (i.e., events that affected the self primarily through their impact on others). Negative and traumatic events that are directly experienced are generally associated with greater adverse outcomes (e.g., symptoms of depression, PTSD) compared to negative and traumatic events directed at others [41]. Second, in accordance with the A1 PTSD diagnostic criterion (DSM-IV-R, [42]), an event was categorized as a trauma if it was life threatening for the individual or someone else involved. Events that meet the A1 criterion are generally associated with greater PTSD symptoms than non-A1 events [43,44]. Events that were non-life threatening were classified as a hardship.

Neighborhood Conditions. The 11 measures of neighborhood conditions are standardized across NGs. Based on recent research reporting distinct relations between multiple dimensions of contextual socioeconomic conditions and outcomes, neighborhood disadvantage and affluence are assessed separately [38,45]. Principal components factor analysis of 2000 Census NG-level measures yielded two factors [46]. The neighborhood disadvantage factor loads negatively on high family incomes, and positively on low family incomes, high levels of poverty, public assistance, unemployment, and vacant housing. The affluence factor loads positively on measures of the proportion of employed civilians ages 16 and over in professional or managerial occupations, the proportion of individuals ages 25 and over who have completed 16 or more years of education, and median home values. Disadvantage and affluence capture distinct but correlated aspects of neighborhood socioeconomic status, much as income and education represent distinct aspects of individual socioeconomic status. Many neighborhoods low in disadvantage are also low in affluence [47]. The traditional family structure scale is composed of percent of families headed by females (reverse-coded), and percent of residents married. The physical disorder measure used here comes from the Systematic Social Observation (SSO) [48,49] component of the CCHAS, in which trained raters assessed 9 ecological conditions (i.e. graffiti, litter, abandoned cars, broken glass) using standardized criteria.

The perceived measures are aggregated to the neighborhood level from individual responses to the Community Survey section of the CCAHS. Scales are based on the PHDCN [50] and use approximately five questions for each scale. Aggregation of reports
from all respondents within the NC minimizes the importance of each respondent’s response, so that the respondent’s own perceptions of neighborhood will not overly influence results [11]. To aggregate individual item responses to factors at the neighborhood level, empirical Bayesian hierarchical linear models were used [51]. The procedure adjusts for missing items, controls for individual socioeconomic characteristics, and improves neighborhood-level estimates by borrowing information from across locations [51]. A perceived disorder measure was also studied because both measures are widely used; social stigma can amplify perceptions [52]. The “collective efficacy” scale includes 10 items from two subscales and assesses a shared willingness (social cohesion) to take action to enforce collective norms (social control) [53]. The total victimization scale captures experiences of being a victim of crime (i.e. assault, property theft, robbery) in the neighborhood. Anomie assesses the extent to which residents report a disconnection from basic societal rules (e.g. agreeing with “Laws are made to be broken”). Reciprocal exchange focuses on the exchange of favors, advice, material goods, and information which make up a social support network within the community [54]. Sociodemographics. Previous research has shown that individual sociodemographic factors predict both the experience of negative life events [2] and the kinds of neighborhoods in which individuals live [55]. We therefore adjust for individual sociodemographics: race/ethnicity (non-Hispanic black, Hispanic, non-Hispanic other, with non-Hispanic White as the reference category), variables indicating whether the respondent is female and is a first generation immigrant, and dummy variables for age (30–39, 40–49, 50–59, 60–69, and 70 years and over, with 18–29 as the reference group), education (12–15 years or 16+ years, with 0–11 years as the reference category), and annual income (of respondent and their spouse if applicable) (less than $5,000, $15,000–$39,999, $40,000 and over, and missing, with $5,000–$15,000 as the reference category). Categories are used for age in order to highlight non-linearity.

**Health Outcomes.** Three health outcomes were selected as commonly studied with respect to both life events and neighborhoods. Descriptions of the depression (CESD) [56], anxiety [57], and self-rated health [58] measures used are given elsewhere.

**Analytic Plan.** The analysis was organized in several stages. First, the prevalence of exposure to each life event in the five years prior to the interview is reported, along with the level of neighborhood similarity of exposure to each event. Neighborhood (i.e. NC) similarity is assessed using intraclass correlations (ICCs), where the classes are neighborhoods. ICCs are reported for each life event category, without and with adjustment for sociodemographics. This method partitions variance in multilevel models attributable to individual versus neighborhood levels as a ratio of the between-neighborhood variance to the sum of the between-neighborhood variance and the between-individual, within-neighborhood variance. The ICC for logit models is calculated as $t/(t+3.29)$, where $t$ is the variance of the random intercepts [59]. ICCs can only be compared within a given dataset, and comparison with other findings from the CCAHS are discussed. Particularly given the level of statistical complexity in the measure and the low

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### Table 1. Categories of Life Events and Prevalence Rates of Recent Negative Life Events Ranked by Level of Neighborhood Similarity.

| Event | Category | Directed at Respondent | Directed at Other | Trauma | Hardship | Prevalence (%) | Unadjusted | Sociodemographic |
|-------|----------|------------------------|--------------------|--------|----------|---------------|-------------|-----------------|
| Lost Job, Someone in HH | X | X | 2.0 | 0.181 | 0.171 |
| Life-Threatening Illness/Accident of Spouse/Child | X | X | 0.8 | 0.132 | 0.066 |
| Death of Child | X | X | 4.8 | 0.132 | 0.044 |
| Moved to Worse Neighborhood | X | X | 9.8 | 0.103 | 0.087 |
| Physical Assault | X | X | 3.2 | 0.098 | 0.050 |
| Serious Financial Difficulties | X | X | 23.0 | 0.082 | 0.030 |
| Robbed/Burglarized | X | - | - | 12.1 | 0.080 | 0.077 |
| Life-Threatening Illness/Accident of Respondent | X | X | 5.7 | 0.077 | 0.061 |
| Legal Trouble | - | - | X | 16.6 | 0.076 | 0.061 |
| Unemployed >3 Months | X | X | 16.4 | 0.074 | 0.032 |
| Unemployed >3 Months, Someone in HH | X | X | 11.5 | 0.059 | 0.043 |
| Divorced | X | X | 4.8 | 0.040 | 0.048 |
| Lost Job | X | X | 13.7 | 0.040 | 0.033 |
| Life-Threatening Illness/Accident of Someone Else Close | X | X | 36.0 | 0.038 | 0.026 |
| Death of Someone Close | X | X | 4.7 | 0.025 | 0.034 |

* Results were adjusted for socio-demographic variables including race/ethnicity, gender, first generation immigrant status, age, education, and annual income. CCAHS, 2001–03. doi:10.1371/journal.pone.0093539.t001
prevalence of some events, the ICCs should be treated as general indicators of the magnitude of clustering rather than as absolute ranks.

Second, a binary indicator of whether the respondent has experienced a recent life event is aggregated to the neighborhood level using the method previously described, and the results mapped. Third, population-weighted regression models assess associations between individual-level characteristics (e.g. socioeconomic status) and life event categories. Fourth, mixed-effects models are estimated to assess the relationship between neighborhood characteristics and life event categories. Mixed-effects models are necessary because multiple respondents living in each neighborhood violates the regression independence assumption. These population-average robust models with population weights include controls for individual socio-demographics not shown. Fifth, to check for sensitivity to neighborhood spatial boundaries, the association between life events and disadvantage is assessed at various spatial scales. Finally, path analyses of the associations among 2 of the most frequently studied [60] neighborhood characteristics (disadvantage and perceived disorder), any recent life event, and 3 health outcomes (self-rated health, depression, and anxiety) are presented. These population-weighted models were estimated using Mplus software [61], with controls for individual sociodemographics not shown.

Results

Overall, 72.7% of the sample experienced at least one negative life event in the last 5 years. Table 1 reports the prevalence of each life event in Chicago in 2001–3. Frequencies of individual life events varied widely across event type, with some events occurring among less than 1% of the sample (e.g. life-threatening illness/accident of a spouse/child) and other events occurring among more than 20% of the sample (e.g. financial difficulties, life-threatening illness of someone else close). For each event in Table 1, ICCs are reported with and without adjustment for covariates. Previous reports of geographic clustering using the same dataset (CCAHS) indicate that most measures of neighborhood physical and social conditions have ICCs above 0.13, some negative emotions (e.g. pessimism) have ICCs of 0.06–0.09, while some physical health and psychological measures have ICCs under 0.05 [13,62]. Thus, the level of clustering of job loss without adjustment for covariates (0.18) is substantial, almost reaching the range of purely neighborhood-based characteristics such as perceived violence in the neighborhood (0.23). The clustering of death of child (0.13) and life-threatening illness/accident to spouse/child (0.13) approximates that of neighborhood-level variables with stronger links to individual characteristics, such as perceived social cohesion (0.14). A moderately high level of clustering was observed for another group of life events, including a move to a worse neighborhood, assault, financial difficulties, robbery/burglary, life-threatening illness/accident of respondent, legal trouble, and unemployment of respondent or household member), which may be consistent with a high degree of clustering of factors at the individual level, and is comparable to the level of geographic clustering previously reported in the same dataset (CCAHS) for negative emotions [13], drinking [63], and exercise [63]. Finally, events including divorce, personal job loss, life-threatening illness/accident of someone else close, and death of someone close, have lower clustering, similar to rates of clustering observed in the CCAHS for measures such as smoking [63], systolic blood pressure, and hopelessness [13].

To assess whether the comparatively high level of clustering observed for negative life events is primarily driven by the tendency of similar individuals and households to live near each, Table 1 also presents ICCs adjusted for sociodemographic variables. In Chicago much of the demographic sorting is due to racial/ethnic and income segregation. Thus, the difference in the levels of clustering between the unadjusted and adjusted ICCs gauges the contribution of segregation and other sorting by neighborhoods to similarity in the experience of negative life events. When adjustment produces larger ICCs, as for divorce and death of someone close, they are interpreted as resulting from disparities running counter to neighborhood selection (e.g. when Whites report greater fear of crime and also live in safer neighborhoods [65]). Overall, our results indicate that clustering of events persists even when adjusted for sociodemographic clustering, suggesting that life events are spatially independent in their own right.

To visually illustrate the extent to which negative life events are geographically clustered, Figure 3 presents a map of the empirical Bayes estimates of the neighborhood prevalence of any recent life event. The highest quartiles of prevalence are colored darker. Consistent with our other results, life events are not randomly distributed with respect to geography. Rather, a clear spatial distribution is discernible such that recent life events are more common in more disadvantaged areas of the city, as well as downtown. Chicago is bounded by Lake Michigan on the east, and the central business district is centrally located on the shore. Respondents living in the outlying areas of the city (distant from downtown, with more single-family houses and wealthier residents) have experienced few recent life events, adjusted for sociodemographics. These are the same outlying areas previously identified as having high levels of social capital [64]. Meanwhile, the central business district (“The Loop”) area also has a comparatively high prevalence of negative life events, despite the low disadvantage of its residents.

To facilitate comparison with prior research, Table 2 reports associations of socio-demographic variables with life event categories. Consistent with previous research [2,25,65], men were significantly more likely to report respondent-directed events, hardships, or any negative life event. Compared to non-Hispanic Whites, non-Hispanic Blacks were more likely to report respondent-directed events and hardships, and Hispanics were more likely to report hardships. Compared to third generation immigrants, first generation immigrants reported far fewer events across all categories, and second generation immigrants reported fewer events overall and respondent-directed events. Older adults also reported having experienced fewer events in the last five years, except for traumas, consistent with findings that negative life events are more common early in the life course [3,65,66]. Having moderate ($30,000–$50,000) and higher ($50,000 or higher) income also predicted having recently experienced fewer negative life events, compared to those in the lower income group ($10,000–$30,000), while the very low income group (often the young and the old) was not significantly different from the lower or higher group (not shown).

Given the high levels of clustering reported in Table 2, the next analytic phase focuses on the relation between negative life events and conventionally studied neighborhood measures. Table 3 examines associations of specific features of neighborhoods with life event categories. Several neighborhood conditions (both objective and perceived) are strong predictors of any recent life event. Certain types of events, specifically hardships and respondent-oriented events tended to be related to the neighborhood measures. The types of events which were significantly associated with neighborhood predictors also tended to be those which were more geographically correlated. Respondents do not necessarily
live in the same neighborhoods as those emotionally close to them, consistent with there being few significant associations of neighborhood characteristics and events to others. Also, many of the traumatic events are related to illness, which is among the more spatially diffuse phenomena studied here. Residents of disadvantaged neighborhoods have greater odds of having recently experienced any event (except other-oriented events). However, affluence is not associated with the risk of negative life events, consistent with a literature showing distinctive findings for disadvantage and affluence.

Other social composition measures also predict life events: having a greater proportion of older adults in the neighborhood is associated with lower odds of respondent-oriented events, hardships, and overall events. Perhaps neighborhood family composition is related to the likelihood of reporting an event to an acquaintance if non-traditional family composition increases the number of people viewed as family. This finding is also consistent with a protective effect of traditional family contexts. Conditions linked with crime (robbery rate and total victimization) and disorder (objective or perceived) strongly predict greater odds of an event directed at the respondent, hardships, and any event, with weaker or no relationships with events directed at others and traumatic events (probably because of the importance of health in this category). The three measures of neighborhood norms and behaviors show divergent patterns of prediction. Anomic increases odds of events directed at the respondent, but not hardships. Collective efficacy decreases odds of respondent-oriented events, hardships, and any event, while reciprocal exchanges between neighbors are non-significant, lending support to the literature on the benefits of collective efficacy beyond other forms of social interaction.

Several studies have found differences in neighborhood “effects” as a function of how neighborhoods are defined [67]. As a cross-check, additional models (not shown) were estimated in which disadvantage and affluence were assessed in concentric rings (“buffers”) around the respondents’ geocoded home addresses with diameters of 250 m, 500 m, 1 km, 1.5 km, 2 km, and 3 km. At each spatial scale, disadvantage coefficients in Poisson models predicting the count of recent life events, adjusting for sociodemographics, were nearly identical (1.1<β<1.2; p<.001). Affluence remained non-significant across spatial scale in the buffer models. This indicates that neighborhood definition does not drive results.

Finally, illustrative results from path analyses of neighborhood characteristics (disadvantage and perceived disorder), life events (any recent event), and 3 health measures (worse self-rated health, depression, anxiety) are given in Table 4. In each model, the

![Figure 3. Neighborhood Estimated Experience of a Recent Life Event (Empirical Bayes Estimates; CCAHS, 2001–03).](doi:10.1371/journal.pone.0093539.g003)
neighborhood characteristic strongly predicts a recent life event, and a recent life event strongly predicts the health outcome. By contrast, direct associations of the neighborhood variable with health are smaller. Analogous results appeared in models on the subsample of respondents who had lived in their neighborhoods for 5 years or more (not shown). The highly significant Sobel test statistics (which tests the indirect effect of neighborhood conditions on health outcomes through recent negative life events) indicate that recent life events partially mediate the associations of neighborhood conditions and health outcomes.

Discussion

These findings illustrate the importance of considering acute rather than only chronic pathways between context and health. Experience of negative life events – to varying extents – are patterned by residential neighborhood context. Previous research on the prevalence of stressful and traumatic life events has tended to consider individuals as the unit of analysis. The evidence in this paper suggests a new direction linking research in a variety of fields (criminology, social networks, urban geography) with the literature on neighborhood health and health disparities. Exposure to actual negative life events is a distinct neighborhood-health pathway beyond fear of events, and may serve as an important mechanism by which context influences stress levels.

Several features of the analysis limit the ability to draw causal links between neighborhoods and events, which is not our goal. Cross-sectional models cannot directly assess risk of event occurrence in neighborhoods. The survey questions do not specify where the events took place. Future data collection may be able to more explicitly capture how events relate to dynamic spatial and social contexts. Respondents who experience negative events closely linked to place, such as muggings, may be more likely to move to places where they feel safer or where they have no unpleasant memories. Or, negative life events may lead to downward socioeconomic mobility, resulting in residence in less desirable areas. These issues do not detract from our aim to demonstrate the association between recent exposure to negative life events and contextual conditions. Future prospective studies could help establish any dynamic longitudinal relationship between residential attainment/choice and life events.

Table 2. Associations of Individual-level Socio-demographic Variables with Negative Life Events (Odds Ratios).

|                  | Directed at Respondent | Directed at Other | Trauma | Hardship | Any Event |
|------------------|------------------------|-------------------|--------|----------|-----------|
|                  | OR                     | OR                | OR     | OR       | OR        |
| Female           | 0.79                   | *                 | 0.92   | 0.86     | 0.77 ***  | 0.79 *    |
| Race/Ethnicity   |                        |                   |        |          |           |
| (ref = Non-Hisp White) |                |                   |        |          |           |
| Non-Hisp. Black  | 1.28                   | *                 | 0.97   | 0.84     | 1.51 ***  | 1.09      |
| Hispanic         | 1.32                   | +                 | 1.04   | 0.92     | 1.51 *    | 1.05      |
| Non-Hisp. Other  | 1.33                   |                   | 0.78   | 0.93     | 1.01      | 1.24      |
| Immigration      |                        |                   |        |          |           |
| (ref = 3rd+ Generation) |            |                   |        |          |           |
| 1st Gen.         | 0.36                   | ***               | 0.49   | ***      | 0.51 ***  | 0.35 ***  | 0.39 *** |
| 2nd Gen.         | 0.69                   | *                 | 0.79   | 0.94     | 0.69 +    | 0.70 *    |
| Age (ref = 18–29) |                        |                   |        |          |           |
| Age 30–39        | 0.89                   |                   | 0.99   | 1.07     | 0.92      | 0.97      |
| Age 40–49        | 0.77                   | +                 | 0.99   | 1.07     | 0.90      | 0.85      |
| Age 50–59        | 0.65                   | *                 | 0.97   | 1.10     | 0.85      | 0.83      |
| Age 60–69        | 0.28                   | ***               | 0.73   | +        | 0.85      | 0.21 ***  | 0.37 *** |
| Age 70+          | 0.18                   | ***               | 0.59   | **       | 0.71 +    | 0.14 ***  | 0.29 *** |
| Education        |                        |                   |        |          |           |
| (ref = 16+ Years) |                        |                   |        |          |           |
| <12 Years        | 1.22                   |                   | 1.10   | 1.16     | 1.08      | 1.17      |
| 12–15 Years      | 1.01                   |                   | 1.11   | 1.05     | 0.94      | 1.16      |
| Income           |                        |                   |        |          |           |
| (ref = $10K-LT$30K) |                    |                   |        |          |           |
| LT$10K           | 1.06                   |                   | 1.02   | 1.16     | 1.09      | 1.02      |
| $30K-LT$50K      | 0.67                   | **                | 1.04   | 1.01     | 0.64 **   | 0.78      |
| $50K+            | 0.62                   | ***               | 1.00   | 1.03     | 0.46 ***  | 0.72 *    |
| Constant         | 3.07                   | ***               | 1.26   | 1.43 +   | 1.87 ***  | 6.03 ***  |

*** p<.001, ** p<.01, * p<.05, + p<.1 (two-tailed tests); CCAHS, 2001–03.
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It is also important to note that respondents are not equally exposed to all events, and not all potential negative events are measured. For instance, divorced respondents must have first been married, and context is associated with both divorce and marriage. Contextual conditions are associated with not only occurrence of an event conditional on eligibility but also eligibility itself. This is because we are interested in the population-level relationship between neighborhood conditions and occurrence of an event. While the structured interview questions covered a broad range of events, many other stressful and potentially traumatic life events, as well as a range of positive events, may be consequential. However, when prompted with an open-ended question, respondents reported relatively few events not previously queried.

### Table 3. Associations of Neighborhood Conditions with Any Life Event in Category, Neighborhood Clusters.

| Objective Measures | Directed at Respondent | Directed at Other | Trauma | Hardship | All Recent Life Events |
|--------------------|------------------------|-------------------|--------|----------|-----------------------|
|                    | Coef. | SE  | p   | Coef. | SE  | p   | Coef. | SE  | p   | Coef. | SE  | p   |
| Disadvantage       | 1.193 | 0.035 | *** | 1.054 | 0.030 | +  | 1.062 | 0.028 | *  | 1.176 | 0.041 | *** | 1.191 | 0.037 | *** |
| Affluence           | 0.980 | 0.023 |  | 0.983 | 0.019 |  | 0.988 | 0.020 |  | 0.970 | 0.024 |  | 0.986 | 0.023 |  |
| Nuclear Family Structure | 0.853 | 0.029 | *** | 0.937 | 0.028 | *  | 0.947 | 0.026 | *  | 0.855 | 0.031 | *** | 0.847 | 0.029 | *** |
| % 65 or Older       | 0.915 | 0.021 | *** | 1.021 | 0.018 |  | 1.002 | 0.018 |  | 0.944 | 0.024 | *  | 0.936 | 0.025 | *  |
| Robbery Rate        | 1.171 | 0.033 | *** | 1.013 | 0.025 |  | 1.039 | 0.023 | +  | 1.143 | 0.035 | *** | 1.160 | 0.033 | *** |
| Disorder            | 1.148 | 0.029 | *** | 1.022 | 0.021 |  | 1.043 | 0.022 | +  | 1.130 | 0.029 | *** | 1.132 | 0.029 | *** |

| Perceived Measures  | Directed at Respondent | Directed at Other | Trauma | Hardship | All Recent Life Events |
|---------------------|------------------------|-------------------|--------|----------|-----------------------|
|                     | Coef. | SE  | p   | Coef. | SE  | p   | Coef. | SE  | p   | Coef. | SE  | p   |
| Disorder            | 1.181 | 0.033 | *** | 1.041 | 0.023 | +  | 1.055 | 0.025 | *  | 1.150 | 0.036 | *** | 1.161 | 0.033 | *** |
| Total Victimization | 1.096 | 0.024 | *** | 1.033 | 0.019 | +  | 1.049 | 0.020 | *  | 1.053 | 0.025 | *  | 1.103 | 0.025 | *** |
| Anomie              | 1.062 | 0.025 | **  | 0.975 | 0.018 |  | 0.996 | 0.018 |  | 1.039 | 0.027 |  | 1.045 | 0.025 | +  |
| Collective Efficacy | 0.909 | 0.024 | *** | 1.002 | 0.019 |  | 0.982 | 0.020 |  | 0.925 | 0.026 | **  | 0.920 | 0.023 | **  |
| Reciprocal Exchange | 0.974 | 0.027 |  | 1.013 | 0.020 |  | 0.998 | 0.019 |  | 0.972 | 0.026 |  | 0.986 | 0.027 |  |
| Unadjusted ICC      | 0.040 | 0.002 |  | 0.012 | 0.000 |  | 0.001 | 0.013 |  | 0.017 | 0.000 |  |

Neighborhood predictors were standardized. Controls for sociodemographics are not shown. Lower ICCs indicate a greater proportion of neighborhood-level variance explained by the predictors.

*** p < .001, ** p < .01, * p < .05, + p > .1 (two-tailed tests).

CCAHS, 2001–03.

Results from separate regressions for each neighborhood-level variable.

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### Table 4. Path Analyses of Neighborhood Characteristics, Life Events, and Health Outcomes, Adjusted for Socio-demographics, CCAHS, 2001–03.

| Neighborhood Predictor | Life Event Mediator | Health Outcome | Neigh->LE | Neigh->Health | LE->Health | Sobel Test |
|-----------------------|---------------------|----------------|-----------|---------------|------------|------------|
|                      |                     |                | Coeff. | p   | Coeff. | p   | Coeff. | p   | Coeff. | p   |
| Disadvantage         | Any Recent          | Self-Rated Health (Worse) | 0.281 | *** | 0.050 | +  | 0.182 | *** | 2.856 | *** |
| Disadvantage         | Any Recent          | Depression     | 0.280 | *** | 0.037 | *  | 0.218 | *** | 3.474 | *** |
| Disadvantage         | Any Recent          | Anxiety        | 0.282 | *** | 0.040 | *  | 0.155 | *** | 3.180 | *** |
| Perc. Disorder       | Any Recent          | Self-Rated Health (Worse) | 0.260 | *** | 0.090 | *  | 0.354 | *** | 3.175 | *** |
| Perc. Disorder       | Any Recent          | Depression     | 0.262 | *** | 0.031 | *  | 0.216 | *** | 4.374 | *** |
| Perc. Disorder       | Any Recent          | Anxiety        | 0.261 | *** | 0.037 | ** | 0.152 | *** | 3.766 | *** |

Neigh = Neighborhood predictor. LE = Life event.

*** p < .001, ** p < .01, * p < .05, + p > .1 (two-tailed tests).

CCAHS, 2001–03.

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Our goal is to quantitatively demonstrate the potential for embedded and interactive experience of conventionally-measured negative life events in risky and resourceful places. Extensive prior ethnographic work has previously described the cumulative nature of life experience burden in poor communities [69]. Whereas ethnographic research has helped elucidate stress pathways in individuals and specific communities, spatial demographic work can provide precedents necessary for comparison across space. The ability to consider spatial variation in life events is required in order to investigate how geographic factors cause life events, how spatial clustering of life events may result in biased estimation of spatial effects, and how life events may lead to further events for those nearby. The results also show that some life events are spatially clustered in ways not driven by poverty, and future work is needed to understand the reason.

Both chronic strain and acute life events may influence physical and mental health over the life course. Preventive efforts may benefit from a careful discourse distinguishing between the impact of fear of events and objective experience of events on health. The work presented here lays a foundation of basic findings on which to bridge research concerning trauma with the literature on neighborhoods and health. We urge stress researchers to consider the embedded nature of specific events within spatial and social network contexts.

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Author Contributions

Conceived and designed the experiments: KK CO. Analyzed the data: KK. Wrote the paper: KK CO.

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