Background:
Genetic components controlling for echocardiographically determined left ventricular (LV) mass are still unclear in the Chinese population.

Methods:
We conducted a family study from the Chin-San community, Taiwan, and a total of 368 families, 1145 subjects, were recruited to undergo echocardiography to measure LV mass. Commingling analysis, familial correlation, and complex segregation analysis were applied to detect component distributions and the mode of inheritance.

Results:
The two-component distribution model was the best-fitting model to describe the distribution of LV mass. The highest familial correlation coefficients were mother-son (0.379, \( P < .0001 \)) and father-son (0.356, \( P < .0001 \)). Genetic heritability (\( h^2 \)) of LV mass was estimated as 0.268 ± 0.061 (\( P < .0001 \)); it decreased to 0.153 ± 0.052 (\( P = .0009 \)) after systolic blood pressure adjustment. Major gene effects with polygenic components were the best-fitting model to explain the inheritance mode of LV mass. The estimated allele frequency of the gene was 0.089.

Conclusion:
There were significant familial correlations, heritability and a major gene effect on LV mass in the population-based families.

Background
Left ventricular (LV) hypertrophy profoundly affects morbidity and mortality from cardiovascular diseases, including myocardial infarction, congestive heart failure, and stroke [1,2]. Therefore, it is important to measure LV mass and manage LV hypertrophy among the hypertensive population in clinical practice [3]. The noninvasive imaging methods of echocardiography have greatly expanded our capability in evaluating increased LV mass, and have thus enhanced our understanding of the natural history of LV hypertrophy [4-6]. Blood pressure and volume overload are recognized as a strong determinant of LV hypertrophy [7]; however, interindividual variations in LV mass can be explained, only to a limited extent, by hemodynamic load [8]. Moreover, LV hypertrophy may occur in the absence of hypertension [9]. Genetic components have been considered as important factors for LV mass, and twin studies have shown that heritability accounts for up to 20–70% [10-13].
The family member study design is a useful tool for investigating genetic and environmental components of complex traits [14]. Significant heritability and evidence of strong familial aggregation in LV mass have been reported in different populations, such as in Caucasian [10,13,15-18], African Americans [17,19], American Indians [12] and Caribbean Hispanic families [20]. But data based on Chinese are not available. Furthermore, investigation on the mode of inheritance by segregation analysis among families can provide parametric estimates for linkage analyses [21]. Even though some major susceptibility genes have already been identified, segregation analysis can provide further evidence for major gene effects in the trait [22]. Previous segregation study showed the mode of inheritance of LV mass was compatible to polygenic model [13], but the study did not prove major gene effects controlling LV mass. Segregation analysis can investigate if the major gene effects exist, besides the polygenic background effects, and help to facilitate the further genomic study.

The study of young families from one community could provide a feasible way to investigate genetic components among the general population, because hypertension complications are comparatively rare and the environmental factors are more homogeneous than hospital-based subjects. The researchers in the Tecumseh Offspring Study showed that among the young population (mean 17 years old), parental LV mass explained a meaningful small percentage for adolescent offspring LV mass variance [23]. The community-based family study was to investigate the mode of inheritance of LV mass in the Taiwanese population.

Methods

Subjects

This family study was part of the Chin-Shan Community Cardiovascular Cohort Study, a prospective cohort study beginning in 1990 [24,25]. The family study arm started in 1997 and was designed to recruit adolescent probands from students in the only junior high school in the community. The proposal was approved by the Institutional Review Board of National Taiwan University Hospital and oral informed consent was obtained in each participant.

At first, a total of 1063 students (with a response rate of 94.6%) agreed to participate in a general health check-up after informed consent was obtained. They underwent examinations for anthropometric measures, blood pressure, lipid profiles and echocardiographic measures. The selected youths were stratified into two groups on the basis of seven measures, including total cholesterol, triglyceride, low density lipoprotein (LDL) cholesterol, body mass index, systolic pressure, diastolic pressure, and high density lipoprotein (HDL) cholesterol. There were 368 students selected for this study. All the students with values below the 10th percentile for HDL cholesterol and above the 90th percentile (for all students in this study) for any of the other six measures, were ascertained as high-risk probands (n = 171). The control young probands (n = 197) were randomly sampling from other young students. After obtaining informed consent from probands' family members, the same measures were performed for each family member.

Because the original stratification was not based exclusively on LV mass and the results of segregation analysis for the families were similar, the results of all families together with correction for ascertainment on proband's status were reported in this study. The ascertainment strategy is to select probands in upper 90th percentiles of the risk factor distribution, and this strategy can substantially increase the power over random sampling. We used the proband high risk status as a surrogate strategy, which could identify individuals in high extreme LV mass distribution in the regressive model.

Echocardiographic examination

Echocardiographic measures have been described in detail previously [4]. Four cardiologists performed and read the measurements. All subjects were asked to lie in the left lateral decubitus position to assure standardized measurement. We recorded the real-time image in video tapes simultaneously and calculated the measures in consecutive beats. We checked the agreement and reliability among these 4 cardiologists as follows. A sub-sample of 15 participants was selected to receive echocardiography by all four physicians to estimate the inter-observer variability of measurement parameters. The intraclass correlation reliability was calculated by a simple replication one-way analysis of variance test [26]. The values of intrarater correlation coefficient reliability of LV mass were 0.80, indicating good agreement.

The LV mass was calculated by the Penn convention [27], and normalized LV mass values were calculated by one allometric value of height of power, 2.7, as suggested by de Simone et al [28].

Statistical analysis

LV mass and related characteristics of the study participants were specified by proband, sibling and parent status. Mean, standard deviation, skewness and kurtosis were presented to show the distribution of related traits. The residual LV mass from the multiple linear regression model, after adjusting for age, gender, body mass index, smoking, and alcohol drinking status, plus sample mean, were used for further genetic analysis.
Genetic analysis

We used commingling analysis on LV mass values to test whether the data were best described by one, two, or three more Normal distributions by ADMIX program [29]. The parameters for each component’s mean, variance and proportion were estimated by the maximum likelihood method. The best-fitted commingling distribution model was defined by comparing the likelihood ratio test statistics among different nested models.

The intra-familial correlation coefficients of LV mass trait were measured in different pairs, including spouse, parent-offspring, and siblings, by FCOR program in S.A.G.E. [30]. We tested if the correlation coefficients of parent-offspring pairs were statistically different from zero by the Fisher’s z test [31]. Heritability estimate of LV mass in the families was estimated by the variance component model, which was implemented in the SOLAR software [32].

Complex segregation analysis

Segregation analysis of adjusted LV mass was conducted using regressive models as implemented in the REGC program in S.A.G.E. [30]. These models assume that the variation of LV mass among family members is the result of a major gene effect, with residual variation reflecting both familial correlations and individual variation. The presence of a major gene is assessed by allowing two factors or alleles (A and B) at a single locus, resulting in three 'ousiotypes' (AA, AB, BB) in individuals. The means of LV mass for each ousiotype is denoted \( \mu_{AA}, \mu_{AB}, \mu_{BB} \), with one common variance of \( \sigma^2 \). The frequencies of allele A and B are denoted \( q_A \) and \( (1- q_A) \), respectively. The distribution of types in the population is assumed to be in Hardy-Weinberg equilibrium. Individuals of each type are assumed to transmit allele A to their offspring with transmission probabilities \( \tau_{AA}, \tau_{AB} \) and \( \tau_{BB} \), respectively. Residual familial resemblance unable to be explained by this major locus is modeled by familial correlations. The correlation between spouses, parents and offspring, mother and offspring, father and offspring, and between siblings are denoted \( \rho_{MM}, \rho_{PO}, \rho_{MO}, \rho_{PO}, \) and \( \rho_{SS} \) respectively. For this study, we adopted class D regressive models, in which residual sib-sib correlations are equal among all sibs of common parentage and can be due to any cause. If \( \rho_{PO} \) is held equal to \( \rho_{SS} \), these models have been shown to be mathematically and numerically equivalent to the conventional mixed model of inheritance in nuclear families [33].

The analyses started with fitting a general model, in which all parameters were allowed to be estimated. Then we compared the general model with various submodels, in which certain parameters were restricted to specific values. Under a Mendelian model, the transmission probabilities, i.e., \( \tau_{AA}, \tau_{AB} \) and \( \tau_{BB} \), were held equal to Mendelian expectations of 1, 0.5, and 0. A nontransmitted environmental effect model predicted that the probability of an individual having one ousiotype or another was independent of both the person’s generation and the ousiotypes of his/her parents. For the environmental model in this study, each of the transmission probabilities was taken to be equal to the factor frequency, i.e., \( \tau_{AA} = \tau_{AB} = \tau_{BB} = q_A \). Both the Mendelian and environmental models can allow for residual familial correlations. A pure polygenic model assumed no major gene effect, so gene frequency and transmission probabilities were all fixed to one. The fit of hierarchical models was compared with the likelihood ratio test, calculated as -2 of the difference between the ln likelihood of the models being compared. The likelihood ratio value follows a chi-square distribution, with degrees of freedom equal to the difference between the number of parameters estimated. Among nonhierarchical models, the most parsimonious model is that with the lowest values of Akaike’s information criterion (AIC = -2 ln likelihood +2 [number of estimated parameters]) [34].

Results

Description of study participants

The characteristics of LV mass and related factors among probands, siblings, and their parents are presented in Table 1. The parents had higher blood pressure, LV mass, and smoking and drinking rates than the probands and siblings. In addition, the parents had larger standard deviations in BMI, blood pressure and LV mass measures than their offspring. The values of skewness and kurtosis showed that LV mass measures were nearly zero, indicating nearly normal distribution in LV mass. The proportion of the variation in LV mass in this sample explained by gender, age, body mass index, smoking and drinking status was 51%.

Commingling analysis results

Commingling analysis showed that a 2-component, rather than a single-component distribution, was the best-fit model for LV mass variations. The component means, variances, and proportions for the 2-component distribution model were (-0.168, 1.029), (0.661, 1.837) and (85.9%, 14.1%), respectively. The \( \chi^2 \) test for comparing the 2-component with the 3-component distribution was not significant (\( \chi^2 = 5.45, \text{df} = 3, \ p = 0.142 \)), indicating
that the 2-component distribution was the best-fit model. It implied that there were major gene effects. The observed and expected distributions from commingling results were plotted in Figure 1. This shows close approximation of expected to observed distribution.

### Familial correlation and genetic heritability estimates

Significant spouse, parent-offspring and sibling coefficients were found (Table 2), and the highest values were mother-son (0.379, \( P < .0001 \)) and father-son (0.356, \( P < .0001 \)). There were significant spouse and sibling correlation coefficients (0.254, 0.169, both \( P < .0001 \)). Genetic heritability \( (h^2) \) of LV mass was estimated as 0.268 ± 0.061 (\( P < .0001 \)), indicating significant genetic components controlling LV mass trait. After adjusting for systolic pressure, heritability decreased to 0.153 ± 0.052 (\( P = .0009 \)).

### Complex segregation analysis result

We used the complex segregation analysis to detect if major gene effects exist to control LV mass trait. We constructed several models, including sporadic, polygenic, environmental and major gene (Mendel) models to estimate the parameters such as gene frequency, transmission, penetrance and correlations from family data (Table 3). Compared with a general unrestricted model, the chi-square and corresponding \( P \) values were presented. We found that the pure major gene effects model and mixed codominant model were similar with the full model (\( P \) values 0.370 and 0.214, respectively). The pure major gene effects model had the smallest AIC value and was considered as the best-fit model to explain the mode of inheritance of LV mass. Under this major gene effects model, we found that the gene, which controlled high LV mass values, had an allele frequency of 0.098, and the three means of LV mass were 64.4, 45.7 and 33.9 gm/m².

### Table 1: Basic characteristics of participants in the study, specified by generations (n = 1,145)

|                  | Probands (N = 368) | Siblings (N = 333) | Parents (N = 444) |
|------------------|-------------------|-------------------|-------------------|
|                  | Mean   | SD    | Skewness | Kurtosis | Mean   | SD    | Skewness | Kurtosis | Mean   | SD    | Skewness | Kurtosis |
| Age (years)      | 16.45  | 0.97  | 0.33     | 0.34     | 17.83  | 3.43  | 0.59     | 1.05     | 43.65  | 5.74  | 1.35     | 4.65     |
| BMI (kg/m²)      | 20.63  | 3.87  | 1.13     | 1.24     | 20.08  | 3.13  | 1.31     | 4.26     | 24.23  | 3.91  | 1.03     | 4.40     |
| SBP (mmHg)       | 108.86 | 13.02 | 0.34     | -0.34    | 111.65 | 12.90 | 0.36     | 0.27     | 119.13 | 15.12 | 0.82     | 1.28     |
| DBP (mmHg)       | 67.79  | 9.33  | -0.05    | -0.75    | 70.66  | 10.36 | 0.18     | 0.51     | 78.48  | 11.46 | 0.70     | 0.68     |
| LV mass (gm)     | 117.02 | 39.99 | 1.00     | 1.06     | 111.10 | 38.26 | 0.91     | 0.87     | 150.08 | 46.65 | 0.62     | 0.21     |
| LV mass index (gm/m²)² | 31.93 | 9.41  | 1.02     | 1.88     | 30.39 | 8.80  | 0.87     | 1.49     | 41.16 | 11.59 | 0.47     | -0.11    |
| Smoking          | 1 (0.3%)|       |          |          | 35 (10.5%) |       |          |          |
| Alcohol drinking | 1 (0.3%)|       |          |          | 25 (7.5%)  |       |          |          |

Abbreviated: BMI: body mass index, SBP: systolic blood pressure, DBP: diastolic blood pressure, LV: left ventricular

**Figure 1**

Observed (histogram) and expected (linear) distributions of adjusted left ventricular mass in the study subjects. Two thin lines were plotted from 2-component Normal distributions with means as (-0.168, 1.029) and variance as (0.661, 1.837), respectively. The thick line was the summation of two thin lines.

**Table 2: Familial correlation coefficients and significant levels for testing the null hypothesis among the various familial pairs in the studied family members.**

| Pairs            | Numbers | Coefficients | \( P \) value |
|------------------|---------|--------------|---------------|
| Spouse           | 162     | 0.254        | 0.001         |
| Parent-offspring | 889     | 0.294        | 0.000         |
| Sibling          | 427     | 0.169        | 0.000         |
| Grandparent-child| 287     | 0.179        | 0.002         |
| Mother-daughter  | 220     | 0.254        | 0.000         |
| Mother-son       | 252     | 0.379        | 0.000         |
| Father-daughter  | 205     | 0.084        | 0.231         |
| Father-son       | 212     | 0.356        | 0.000         |
| Sister-sister    | 131     | 0.136        | 0.121         |
| Sister-brother   | 209     | 0.162        | 0.019         |
| Brother-brother  | 87      | 0.208        | 0.053         |
Table 3: Parameter estimates from segregation analysis of residual LV mass values: Class D regressive models, conditional on proband phenotypes

| Model | Pure familial correlation, i.e. polygenic | Pure major gene (Mendel) | Mendel + familial correlation (mixed codominant) | A Dominant, $\mu_{AB} = \mu_{AA}$ | A Recessive, i.e. B Dominant, $\mu_{AB} = \mu_{BB}$ | Unrestricted, general |
|-------|----------------------------------------|--------------------------|-----------------------------------------------|-----------------------------------|-----------------------------------------------|----------------------|
|       | Parameter | S.E. | Parameter | S.E. | Parameter | S.E. | Parameter | S.E. | Parameter | S.E. | Parameter | S.E. |
| $q_A$ | [1] | 0.098 | 0.028 | 0.093 | 0.030 | 0.040 | 0.020 | 0.000 | 0.142 | 0.109 | 0.024 |
| $\tau_{AA}$ | - | [1] | [1] | [1] | [1] | [1] | [1] | [1] | 1.000 | 0.000 |
| $\tau_{AB}$ | - | [0.5] | [0.5] | [1] | [1] | [1] | [1] | 0.249 | 0.100 |
| $\tau_{BB}$ | - | [0] | [0] | [0] | [0] | [0] | [0] | 0.034 | 0.030 |
| $\mu_{AA}$ | 36.0 | 0.5 | 64.4 | 4.3 | 65.1 | 4.5 | 51.4 | 2.8 | 40.9 | 8.2 | 62.7 | 3.4 |
| $\mu_{AB} = \mu_{AA}$ | 45.6 | 1.8 | 45.7 | 1.9 | 51.4 | 2.8 | 36.0 | 0.5 | 46.1 | 1.4 |
| $\mu_{BB} = \mu_{AA}$ | 33.7 | 0.5 | 33.9 | 0.6 | 34.9 | 0.6 | 36.0 | 0.5 | 33.5 | 0.6 |
| $\sigma^2$ | 80.6 | 4.6 | 52.4 | 4.2 | 53.5 | 4.7 | 61.8 | 5.5 | 80.6 | 4.6 | 47.1 | 4.9 |
| $\rho_{MF}$ | 0.199 | 0.072 | 0.327 | 0.093 | 0.315 | 0.091 | 0.246 | 0.079 | 0.199 | 0.072 | 0.334 | 0.098 |
| $\rho_{MO}$ | 0.196 | 0.050 | [0] | 0.058 | 0.070 | 0.130 | 0.067 | 0.196 | 0.050 | 0.076 | 0.079 |
| $\rho_{FO}$ | 0.093 | 0.059 | [0] | -0.070 | 0.069 | -0.025 | 0.064 | 0.093 | 0.059 | -0.051 | 0.076 |
| $\rho_{SS}$ | 0.118 | 0.055 | [0] | 0.000 | 0.000 | 0.037 | 0.069 | 0.118 | 0.055 | 0.017 | 0.065 |
| # of parameters | 6 | 6 | 9 | 8 | 8 | 12 |
| $-2\ln(L)$ | 4721.80 | 4684.40 | 4682.38 | 4688.22 | 4721.80 | 4677.90 |
| AIC | 4733.8 | 4696.4 | 4700.38 | 4704.22 | 4737.8 | 4701.9 |
| $\chi^2$ | 43.9 | 6.5 | 4.48 | 10.32 | 43.9 |
| $P$ | 0.000 | 0.370 | 0.214 | 0.035 | 0.000 | Baseline |

Abbreviation: $q_A$, allele frequency; $\tau_{AA}$, $\tau_{AB}$, $\tau_{BB}$, transmission probability; $\mu_{AA}$, $\mu_{AB}$, $\mu_{BB}$, genotype means; $\sigma^2$, variance; $\rho_{MF}$, $\rho_{MO}$, $\rho_{FO}$, $\rho_{SS}$, correlation coefficients among spouse, mother-offspring, father-offspring and sibling pairs; #, number of parameters; $\ln(L)$, logarithm of likelihood; AIC, Akaike information criteria; $P$, p-value by compared with the unrestricted general model.
also demonstrated and there was significant spouse correlation up to 0.327.

Discussion

Significant heritability and familial correlations of LV mass were proven in this study, based on a large number of adolescent proband families. We also demonstrated that the inheritance mode of LV mass was compatible with major genes effects among the Chinese population. This study had two scientific contributions. First, there is considerable homogeneity in this study population. Most of the subjects live in the same community; hence their social and living environments tended to be more similar than those in different communities. Also, the characteristics of study subjects can avoid the potential bias of recruitment from a hospital setting. Second, the results are particularly relevant for a population at low risk for atherosclerosis, since the probands were systematically selected from young adolescents in the community, in which the complication of hypertension and LV hypertrophy were not prevalent in the study population.

There are many reports on left ventricular mass heritability with estimates from different populations (Table 4). Most studies showed significant heritability and sibling correlations, especially among African Americans. Our study results were compatible with Caucasian population studies [13,15,17]. The high heritability estimates in African Americans might be due to ascertainment from hypertensive hyperlipidemic proband siblings [19]. Caribbean Hispanics population had high estimate [20]. Variation on heritability estimates of LV mass might be explained by different ethnicity. Also, the different expression of a complex polygenic trait in relation to unmeasured factors might be involved.

The heritability decreased after adjusting for blood pressure. It implied possible pleiotrophic effects of genes on controlling blood pressure and LV mass. Hemodynamic load, such as stroke volume, has an influence on LV mass among young adults and adolescents, and this impact is more important than body size. However, high proportions of LV mass variations still remain unexplained [8]. Genetic components played important role in residual LV mass variations. We ascertained young proband families, where hypertension complications are comparatively rare, and from one community population, where environmental factors are rather homogeneous for genetic studies. Our study subjects were rather young, and our results showed results similar to the Tecumseh offspring study [23].

We found the highest correlations were in parent-son pairs, which indicated male offspring had influential effects from parents. Our estimate of sibling correlation was as the same as Framingham sibling pairs (0.16), while the estimates of parent-offspring and spouse correlations (0.29 and 0.25, respectively) were much higher than those in Framingham pairs (0.15 in parent-offspring, 0.05 in spouse) [15]. For intra-familial resemblance, greater father-offspring correlations were observed among the British Caucasian population, much higher than mother-child correlations [17]. Our study also showed sex-specific patterns of familial correlations; the highest was the parent-son pairs (0.36–0.38) and the lowest was the father-daughter (0.08). Different parent-of-origin effects on offspring LV mass was reported in European families [37]. The mother-offspring correlation coefficient was significantly higher than father-offspring correlation among the European population. The discrepancy implies the importance of ethnic difference on parental factors on offspring traits. Also, high spouse correlation implies that the common household effects were important for controlling LV mass. Lifestyle habits, such as salt intake and physical activity, might be also familial aggregation and may thus explain the high correlation between spouse pairs.

There were several studies on genomic profiles for LV mass. For example, a cross-breeding hypertensive rat model demonstrated two loci with high LOD scores [38,39]. There were many reports on candidate genes such as G-protein beta-3, aldosterone synthase, and beta-1 adrenoceptor genes associated with LV mass [18,40]. Our study results can provide further genomic research on LV mass.

Study limitations

The limitations of our study were as follows. Firstly, only mathematical modeling methods such as commingling and segregation analyses were investigated and no candidate gene markers were investigated. Although we postulated one major gene with allele frequency around 0.1 controlling LV mass and 2-component commingling patterns, we did not shed light on which genes most likely involved. Secondly, epistasis among genes or gene-environmental interaction cannot be explored in this study. Incorporation with environmental factors can elucidate the possible roles of risk factors and interaction effects.

Conclusion

We showed that there were significant parent-son correlation coefficients, genetic heritability and major gene effects controlling LV mass among ethnic Chinese in Taiwan. Candidate gene markers could be used to investigate the association and linkage with LV mass.

Competing interests

The author(s) declare that they have no competing interests.
Table 4: Summary of heritability estimates of left ventricular mass trait

| Author, year | Study design | Study population origins | Numbers of families/subjects | Adjusted covariate | Heritability estimate ($h^2$) | Comments |
|--------------|--------------|--------------------------|-----------------------------|-------------------|------------------------------|----------|
| Post et al. 1997 | Relative pairs | Caucasians, Framingham families, no systemic diseases | 2624 subjects | Age, height, weight, systolic blood pressure | 0.24–0.32 | Limit in subjects without systemic diseases nor medication history |
| Garner et al. 2000 | Nuclear families | White, European families | 149/624 | weight | 0.28 | Children population |
| Kotchen et al. 2000 | Sibling | African American, hypertensive, hyperlipidemia | 68 sibpairs | Age, gender | 0.65–0.72 | FCOR and ASSOC in SAGE program to perform the analysis |
| Arnett et al. 2001 | Sibling | African Americans & white hypertensive | 1664 siblings | Age, gender, systolic blood pressure, obesity | Sibling correlations 0.29–0.44 in African American, 0.04–0.12 in white | High sibling correlation in African Americans |
| Palatini et al. 2001 | Nuclear families | Tecumseh Offspring Study | 251 offspring and 290 parents | Age, gender, body height systolic blood pressure, insulin, urinary sodium excretion, | No estimate of $h^2$, parental LV mass explained 7.6% of total variance of offspring LV mass | Multiple linear regression models Young offspring (mean 17 years old) |
| Mayosi et al. 2002 | Extended families | British Caucasians, hypertensive probands from hospital or clinics | 229/955 | Age, systolic blood pressure, weight, height, WHR, diabetes | 0.23–0.29 | Ascertainment correction: yes |
| Swan et al. 2003 | Twins | Caucasians, Population-based | 55 pairs MZ vs. 55 pairs DZ | Age, gender, blood pressure, weight | 0.53–0.69 | |
| Bella et al. 2004 | Relative pairs, mostly sibpairs | American Indian families, different geographic location | 455/1373, 1305 relative pairs, 1077 sibpairs | Age, gender, centers, weight, height, systolic blood pressure, heart rate, medication, diabetes | 0.17 (multiple adjusted) - 0.27 (first three variates adjusted) | No significant heritability in Arizona Indians |
| Juo et al. 2005 | Extended families | Caribbean Hispanics | 84/623 | Age, gender, weight | 0.51 | Adding systolic blood pressure, diabetes, medication did not affect the estimate |
| This report | Nuclear families | Ethnic Chinese, young probands, community-based | 368/1145 | Age, gender, body mass index, blood pressure | 0.15 (multiple adjusted) - 0.27 (first 3 variates adjusted) | |
Authors’ contributions
KLC carried out the data collection, statistical analyses, participated in the study design and processing data. YTL participated in the design of the study and supervised the ideas developed in hypothesis generation. MFC & HCH performed the laboratory measurement in lipid levels and were in charge of quality control. YCS carried out the data collection. All authors read and approved the final manuscript.

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