Impact of Economic Policy Uncertainty on Carbon Emissions: Evidence from 137 Multinational Countries

Hai-Jie Wang 1, Yong Geng 2*, Xi-Qiang Xia 1 and Quan-Jing Wang 1,*

1 School of Business, Zhengzhou University, 100 Kexue Road, Gaoxin District, Zhengzhou 450001, China; whj@zzu.edu.cn (H.-J.W.; xqxia@zzu.edu.cn (X.-Q.X.)
2 School of Environment and Science Engineering, Shanghai Jiaotong University, 800 Dongchuan Road, Minhang District, Shanghai 200240, China; ygeng@sjtu.edu.cn

* Correspondence: wqj@zzu.edu.cn

Abstract: With growing economic policy uncertainty (EPU) and the importance of protecting the natural environment worldwide, the relationship between EPU and carbon emissions should be investigated further. However, conclusions in the existing literature on the relationship between EPU and carbon emission are inconclusive. This paper aims to examine the influence of EPU on carbon emissions according to the Stochastic Impacts by Regression on Population, Affluence and Technology (STIRPAT) model. To investigate such essential issues, we conduct GMM estimations by utilizing cross-country data covering 137 countries during the period 1970–2018, obtained from World Bank and OECD statistics. Our empirical estimations support that EPU would bring about more carbon emissions, while we conduct empirical analysis by changing the system of measurement, employing alternative estimation and constructing new samples. Our study provides substantial policy implications for government participation in international treaties on environmental protection to mitigate environmental degradation.

Keywords: economic policy uncertainty; carbon emissions; GMM; multinational research

1. Introduction

Nowadays, due to the outbreak of the COVID-19 pandemic and growing tension concerning the international environment, economic policy uncertainty (EPU) is growing rapidly worldwide (Baker et al., 2020; Jordà et al., 2020; Bakas and Triantafyllou, 2020) [1–3]. EPU can not only change economic activities or the attention of policymakers (Balcilar et al., 2016; Degiannakis et al., 2018; Hailemariam et al., 2019; Phan et al., 2021) [4–7], but also change a firm’s behaviors or decisions in the direction of environmental protection (Guidolin and La Ferrara, 2010; Kang et al., 2017; Olanipekun et al., 2019; Akron et al., 2020) [8–11], since economic activities and the manner of production contribute to air pollution (Salahuddin et al., 2018; Shahbaz et al., 2019) [12,13]. With growing social problems such as human diseases, extreme climate events, and natural disasters caused by serious air pollution (Kompass et al., 2018; Shahbaz et al., 2019; Coskuner et al., 2020) [13–15], it is generally accepted globally that to reduce air pollution is essential for human survival and for national sustainable development (Hambira et al., 2020; Huo et al., 2020) [16,17]. Thus, it is necessary to query whether increasing EPU can affect air pollution (Jiang et al., 2019; Yu et al., 2021; Adams et al., 2020) [18–20].

In fact, a small number of studies have begun to focus on the relationship between EPU and air pollution, especially after the research of Jiang et al., (2019) [18] who declared that EPU can affect direct government policy, which may promote or hinder carbon emissions. Later, scholars began to query the effect of EPU on CO2 emissions. However, there are some gaps in the existing studies. Specifically, the conclusions are controversial (Gill et al., 2019; Shabir et al., 2021) [21,22], and while much of the literature has pointed out that EPU brings more environmentally adverse effects (Yu et al., 2021, Adedoyin and Zakari,
Int. J. Environ. Res. Public Health 2022, 19, 4

2021) [19,23], a small number of scholars have argued that EPU does some good in reducing CO₂ emissions (Chen et al., 2021; Doğan and Güler, 2021) [24,25], and other researchers argue that there is no influence of EPU on CO₂ emissions (Abbasi and Adedoyin, 2021; Liu and Zhang, 2021) [26,27]. Additionally, most studies have utilized the EPU index given by Baker et al., (2016) [28] to measure EPU, but this index has some limitations in that it cannot reflect uncertainty related to political events and it is not calculated from a single base for different countries worldwide, which may bring about biased results (Adams et al., 2020; Atsu and Adams, 2021) [20,29]. Besides, empirical investigations of the impact of EPU on CO₂ emissions are utilizing simple data for one or a few countries (Pirgaip and Dinçergök, 2020; Adedoyin and Zakari, 2021; Sohail et al., 2021) [23,30,31]. Aside from these, some studies have conducted an empirical investigation into this issue by employing the Stochastic Impacts by Regression on Population, Affluence and Technology (STIRPAT) model, which is generally utilized to conduct empirical investigation for issues such as CO₂ emissions (Amin and Dogan, 2021; Anser et al., 2021; Yu et al., 2021) [19,32,33].

This scenario naturally motivates us to raise an interesting issue: what effect does EPU exert on air pollution? To answer this, we conducted an empirical investigation by employing multinational panel data for 137 countries during the period 1970–2018 via GMM estimation according to the STIRPAT model. We measured EPU by the variable World Uncertainty Index provided by Ahir et al., (2018) [34], which can help us to comprehensively examine the role of EPU in air pollution. Compared to the previous literature, the potential contributions of this work are as follows. Firstly, we carry out empirical estimation for the impact of EPU on air pollution based on multinational panel data for 137 countries via the world uncertainty index for EPU provided by Ahir et al., (2018) [34], which can offer further common findings on EPU’s influence on air pollution (Chen et al., 2021) [24]. Secondly, only a few studies have utilized the STIRPAT model, which is generally utilized in order to investigate the influence of EPU on carbon emissions, and in this study we attempted to examine the impact of EPU on carbon emissions by conducting static and dynamic estimations using the STIRPAT model.

The rest of this paper is structured in line with previous empirical studies. Specifically, Section 2 is mainly concerned with the variable data and estimation method. Section 3 offers the main results and discussion. Section 4 gives a brief summary of our empirical results and proposes some suggestions to policymakers.

2. Materials and Methods

2.1. Variables

Air pollution (CO₂): In line with Shao et al., (2011) [35], CO₂ emission is the most critical among environmental pollutants and is generally accepted as an indicator of environmental impact or climate change. Due to the well-established database on CO₂ emission, we also measure CCP by CO₂ emissions (denoted by CO₂) (the unit of CO₂ is one million tons).

Economic policy uncertainty (Uncertainty): variables for EPU such as the EPU index, world trade uncertainty index (WTUI) and world uncertainty index (WUI) were generally utilized in previous studies investigating the issues of EPU (Pirgaip and Dinçergök, 2020; Adams et al., 2020; Qi et al., 2021) [20,30,36]. However, the data for EPU only covers 28 countries, while the data for WTUI and WUI was constructed for 143 countries from 1996 onwards. Besides, compared to the EPU index, the WUI index can capture the uncertainty of political events and is built according to a common base for all countries (Atsu and Adams, 2021) [29]. To come to conclusions on the role of EPU in carbon emissions worldwide, we utilize the WUI data to measure the EPU in line with Adams et al. (2020) [20]. Specifically, Ahir et al. (2018) [34] construct quarterly indices of EPU by using frequency counts of “uncertainty” (and its variants) in the quarterly Economist Intelligence Unit (EIU) country reports, which are scaled by the total number of words in each report and multiplied by 1000 for worldwide comparability. The EPI data is year-frequency, and takes
the mean value of quarterly WUI to measure EPU annually (Atsu and Adams, 2021) [29],
denoted by Uncertainty. A higher value of Uncertainty usually stands for higher EPU.

According to previous articles focusing on carbon emissions and the extended STIR-
PAT model (Wen et al., 2016; Wang et al., 2021b; Yang et al., 2021) [37–39], we include the
following variables in our model:

Economic performance (GDP): As suggested in the traditional STIRPAT model, Afflu-
ence is usually measured by national real income or GDP per capita, so in this study we
also measure Affluence by the national real GDP per capita, which is constant in 2011 US
dollars and denoted by GDP.

Population (POP): In the STIRPAT model, Population is often captured by the total
number of people; to control the influence of population on carbon emissions, we also
utilize this indicator to measure population, denoted by POP (Verma et al., 2021) [40].

Green technologies (GI): In the STIRPAT model, T often stands for technologies which
can produce goods in an environmentally-friendly manner or that increase energy efficiency
which can reduce carbon emissions. In this study, we capture T by green technologies, mea-
sured by the total amount of patent applications in environmental management, denoted
by GI.

Density of people (Density): Aside from the total population, we also measure popula-
tion by density, as higher density may bring about more demand in terms of deforestation
in order to offer more space for human living. As provided by Bottero et al., 2017 [41], pop-
ulation density is calculated by the number of people per square km, denoted by Density.

Structure of population (Aging): Since the aging problem can change a government’s
attitude to natural environmental protection, we measure the aging problem by the share
of people aged 65 or above in relation to total population (Wang et al., 2021a) [42], denoted
by Aging.

Industrial share (IND): To control the influence of industrial structure on carbon
emissions, we measure industrial structures by the proportion of industry value added
(including construction) to GDP (Usman et al., 2021) [43].

Process of urbanization (Urban): As suggested by Dale (2018) [44], we use the share of
urban residents to total population to measure the urbanization rate, which can capture
the potential influence of urbanization on carbon emissions.

International trade (Trade): Similarly to Li et al., (2017) [45], we utilize the share of
exports and imports of goods and services as a proportion of GDP to capture the degree of
openness, denoted by Trade.

Utilization of foreign investment (IFDI): To measure production activities created by
foreign direct investment, we include the control variable of FDI as calculated by the ratio
of net inward FDI to GDP, which is supported by Mahadevan and Sun (2020) [46].

Table 1 provides the details of such variables.

| Variables | Definition | Source |
|-----------|------------|--------|
| Dependent variables | | |
| CO₂ | The total amount of CO₂ emissions | WDI (2020) [47] |
| Uncertainty | The world uncertainty index | Ahir et al., (2018) |
| Independent Variables | | |
| GDP | GDP per capita (constant 2010 US$) | WDI (2020) [47] |
| POP | The number of total population, whose unit is million | WDI (2020) [47] |
| GI | The number of patent applications about the environmental management | OECD (2020) [48] |
| Density | Population density (people per sq. km of land area) | WDI (2020) [47] |
| Aging | Population ages 65 and above (% of total population) | WDI (2020) [47] |
| IND | Industry (including construction), value added (% of GDP) | WDI (2020) [47] |
| Urban | Urban population (% of total population) | WDI (2020) [47] |
| Trade | Share of the sum of exports and imports of goods and services to GDP | WDI (2020) [47] |
| IFDI | Foreign direct investment, net inflows (% of GDP) | WDI (2020) [47] |
2.2. Data

Data on CO$_2$ are sourced from the database of the World Bank. Data on Uncertainty are provided by Ahir et al., (2018) [34] (Ahir et al., (2018) [34], The world uncertainty index, Website of World Uncertainty Index: http://www.policyuncertainty.com/wui_quarterly.html, accessed on 13 November 2021). Data for GI is obtained from the OECD (2020) [48], while data for remaining variables are given by World Development Indicators (WDI (2020) [47] World Bank, World Development Indicators. http://databank.worldbank.org/data/reports.aspx?source=wdi-database-archives-(beta), accessed on 13 November 2021). We merge all data together according to dimensions such as country and year, and then filter the observations by following principles such as dropping observations with missing values. We finally construct a panel dataset for 137 countries from 1970 to 2018. Similar to Wang et al., (2021b) [38], we take the log for these variables.

The basic distribution of the variables is listed in Table 2. For the CO$_2$ variable, it is obvious that the mean for CO$_2$ is 3.133, and the median of is 3.071, while the minimum, standard deviation (S.D), and maximum are 0.047, 1.953, and 9.206, respectively, which means that CO$_2$ varies greatly among different countries. When we analyze the distribution of EPU, we find that the min, mean, median and max of Uncertainty are 0, 0.127, 0.1 and 0.851, respectively, while S.D is 0.112.

Table 2. Summary of descriptive statistics.

| Variable | N    | Mean   | Min   | P25  | Median | P75  | Max   | S. D |
|----------|------|--------|-------|------|--------|------|-------|------|
| CO$_2$   | 4814 | 3.133  | 0.047 | 1.463| 3.071  | 4.569| 9.206 | 1.953|
| Uncertainty | 4814 | 0.127  | 0.000 | 0.044| 0.100  | 0.185| 0.851 | 0.112|
| GDP      | 4814 | 8.201  | 5.306 | 6.979| 8.117  | 9.292| 11.431| 1.501|
| POP      | 4814 | 16.362 | 13.049| 15.377|16.133 | 17.261|21.055|1.374|
| GI       | 4814 | 0.996  | 0.000 | 0.000| 0.000  | 1.386| 8.203 | 1.738|
| Density  | 4814 | 3.986  | 0.748 | 3.076| 4.089  | 4.800| 8.981 | 0.573|
| Aging    | 4814 | 1.891  | 0.522 | 1.438| 1.687  | 2.417| 3.353 | 0.785|
| IND      | 4814 | 3.309  | 0.835 | 3.110| 3.316  | 3.518| 4.486 | 0.583|
| Urban    | 4814 | 3.843  | 1.347 | 3.532| 3.996  | 4.295| 4.615 | 0.581|
| Trade    | 4814 | 4.146  | 0.155 | 3.836| 4.153  | 4.504| 6.095 | 0.575|
| IFDI     | 4814 | 1.088  | 0.000 | 0.449| 1.002  | 1.573| 4.648 | 0.785|

2.3. Estimating Methods

According to the previous literature, such as Wen et al., (2016) [37] and Wang et al., (2021b) [38], carbon emissions are not only influenced by current situations, but are also affected by previous environmental performance; to include this dynamic process into our estimation, we choose the generalized moment method to conduct empirical investigations. However, as suggested by Wang et al., (2019) [49], the GMM estimation system has some advantages over other estimation methods, so we finally utilized the GMM estimation system to conduct our empirical investigation, which is given as following.

$$CO_{2it} = \alpha_1 CO_{2i,t-1} + \theta Uncertainty_{it} + \beta' X + u_i + u_t + \epsilon_{it}$$ (1)

where $i = 1, 2, 3 \ldots N$ stand for the individual country, whereas $t = 1, 2, 3 \ldots T$ represent the dimension of year. CO$_2$ is the dependent variable, whereas CO$_2$$_{2,t-1}$ is the first lag. Uncertainty is the variable of EPU. $X$ represents the control variables, $\beta$ is the vector for the corresponding coefficients, and $u_i$ and $u_t$ stands for the fixed effect of individual country and year, respectively.

3. Results and Discussion

3.1. Baseline Results

At first, we estimate the standard STIRAT model with EPU by using the GMM estimation system, whose results are shown in column (1) of Table 3. The coefficient of uncertainty is 0.138, which is significantly positive even at 1%, indicating that the EPU
would bring about more CO$_2$ emissions. When we take other variables for the extended STIRPAT model into account, the coefficient of uncertainty in column (2)–(4) is 0.171, 0.182, 0.172, respectively, and all are significantly positive at 1%, again confirming the idea that EPU would cause more CO$_2$ emissions.

Table 3. SYS-GMM estimations for EPU’s impact on EPI.

|                  | (1)            | (2)            | (3)            | (4)            |
|------------------|----------------|----------------|----------------|----------------|
| L. CO$_2$       | 0.754 ***      | 0.717 ***      | 0.505 ***      | 0.523 ***      |
|                 | (23.278)       | (17.718)       | (9.612)        | (9.848)        |
| Uncertainty     | 0.138 ***      | 0.171 ***      | 0.182 ***      | 0.172 ***      |
|                 | (4.008)        | (4.208)        | (3.688)        | (3.586)        |
| GDP             | 0.313 ***      | 0.430 ***      | 0.465 ***      | 0.496 ***      |
|                 | (6.034)        | (6.617)        | (5.533)        | (5.746)        |
| POP             | 0.522 ***      | −1.657 ***     | −7.231 ***     | −6.725 ***     |
|                 | (6.768)        | (−3.210)       | (−6.855)       | (−6.604)       |
| GI              | −0.007 ***     | −0.006 ***     | −0.005 ***     | −0.004 *       |
|                 | (−3.168)       | (−3.091)       | (−2.603)       | (−1.949)       |
| Density         | 1.975 ***      | 7.262 ***      | 6.771 ***      |                |
|                 | (3.848)        | (7.103)        | (6.752)        |                |
| Aging           | −0.730 ***     | −1.628 ***     | −1.513 ***     |                |
|                 | (−3.858)       | (−5.329)       | (−5.231)       |                |
| IND             | 0.008          | 0.007          |                |                |
|                 | (0.625)        | (0.432)        |                |                |
| Urban           | 0.354 ***      | 0.276 ***      |                |                |
|                 | (5.945)        | (4.464)        |                |                |
| Trade           | −0.004         | (−0.260)       |                |                |
| IFDI            | −0.001         | (−0.241)       |                |                |
| Year FE         | Yes            | yes            | yes            | yes            |
| N               | 4667           | 4667           | 4667           | 4667           |
| AR (1)          | −4.787         | −4.749         | −3.911         | −4.023         |
| AR (1)-P        | 0.000          | 0.000          | 0.000          | 0.000          |
| AR (2)          | −1.346         | −1.283         | −0.926         | −0.979         |
| AR (2)-P        | 0.178          | 0.200          | 0.354          | 0.327          |
| Hansen-P        | 0.954          | 0.946          | 0.983          | 0.974          |

Notes: *** and * indicate statistical significance at the 1%, and 10% levels, respectively. Z-statistics are in parenthesis.

These GMM estimation results strongly support that EPU exerts a positive impact on CO$_2$ emissions, supporting previous literature such as that of Adams et al., (2020) [20] and Qi et al., (2021) [36]. The potential reason for this is that, once the EPU increases, the more the attention of policymakers would turn to how to maintain economic growth by stimulating economic activities, thus reducing the importance of environmental protection during the implementation of these policies. Furthermore, with increasing EPU, the activity of research and development in clean energy and energy-saving technologies experiences some decline and firms prefer to use cheap fossil fuels (Ulucak and Khan, 2020) [50], which may cause more carbon emissions (Al-Thaqeb and Algharabali, 2019) [51].

3.2. Robustness Test

To prove our idea that EPU positively affects CO$_2$ emissions, we further carried out several robustness tests.

3.2.1. Other Variables of EPU

Firstly, to avoid the potential bias caused by the measurement, we utilize another three variables to capture the EPU, provided by Ahir et al., (2018) [34]. The first is trade EPU, denoted by uncertainty_trade; the second is the absolute count of the word “uncertainty” in the quarterly Economist Intelligence Unit (EIU) country reports, denoted by uncer-
tainty_absolute; while the third is similar to the EPU, but is adjusted by the nearest two seasons, denoted by uncertainty_season. (More detailed information on such variables can be seen in: http://www.policyuncertainty.com/media/WUI_mimeo_10_29.pdf, accessed on 13 November 2021) Table 4 provides the estimation results when we utilize these three variables to measure the EPU. From Table 4 we can obtain that the coefficient of uncertainty_trade is 0.059, which passes the significance test at 5%, suggesting that a higher trade EPU often brings more CO$_2$ emissions. Moreover, the coefficient of uncertainty_absolute and uncertainty_season in column (2) and (3) is 0.023 and 0.083, respectively, offering strong evidence for the conclusion that EPU would lead to more CO$_2$ emissions.

Table 4. Robustness test—changing the measurement of EPU.

|                  | (1)         | (2)        | (3)        |
|------------------|-------------|------------|------------|
| L. CO$_2$        | 0.403 ***   | 0.506 ***  | 0.518 ***  |
|                  | (3.313)     | (8.662)    | (9.758)    |
| Uncertainty_trade| 0.059 **    |            |            |
|                  | (2.244)     |            |            |
| Uncertainty_absolute| 0.023 ***  |            |            |
|                  | (3.390)     |            |            |
| Uncertainty_season|            |            | 0.083 *    |
|                  |             |            | (1.667)    |
| GDP              | 0.983 ***   | 0.451 ***  | 0.514 ***  |
|                  | (5.543)     | (5.259)    | (5.878)    |
| POP              | 10.477 *    | -6.987 *** | -6.327 *** |
|                  | (1.663)     | (-5.943)   | (-6.312)   |
| GI               | 0.002       | -0.005 **  | -0.004 *   |
|                  | (0.608)     | (-2.239)   | (-1.839)   |
| Density          | -9.044      | 7.001 ***  | 6.341 ***  |
|                  | (-1.457)    | (6.162)    | (6.430)    |
| Aging            | -0.453 *    | -1.438 *** | -1.541 *** |
|                  | (-1.658)    | (-4.428)   | (-5.263)   |
| IND              | -0.022      | 0.008      | 0.005      |
|                  | (-0.792)    | (0.483)    | (0.316)    |
| Urban            | -0.095      | 0.375 ***  | 0.282 ***  |
|                  | (-0.460)    | (4.881)    | (4.541)    |
| Trade            | 0.007       | 0.004      | -0.003     |
|                  | (0.537)     | (0.265)    | (-0.208)   |
| IFDI             | -0.005      | -0.002     | -0.001     |
|                  | (-1.198)    | (-0.457)   | (-0.181)   |
| Year FE          | yes         | yes        | yes        |
| N                | 2818        | 4670       | 4667       |
| AR (1)           | -2.964      | -4.095     | -3.908     |
| AR (1)-P         | 0.003       | 0.000      | 0.000      |
| AR (2)           | -0.169      | -0.922     | -0.920     |
| AR (2)-P         | 0.866       | 0.356      | 0.357      |
| Hansen-P         | 0.733       | 0.926      | 0.960      |

Notes: ***, ** and * indicate statistical significance at the 1%, 5%, and 10% levels, respectively. Z-statistics are in parenthesis.

3.2.2. Changing the Measurement of Carbon Emissions

Secondly, we re-conduct the empirical examination by utilizing another three measurements of GHG emissions, total GHG emissions, CH$_4$ emissions and N$_2$O emissions, as suggested by Lin et al., (2017) [52]. We offer the results in Table 5. While we control the time fixed effect, individual fixed effect and other variables which may affect CO$_2$ emissions, the coefficient of uncertainty in column (1) is 0.293, which is significantly positive, again indicating that the EPU would bring about more CO$_2$ emissions. Furthermore, while we capture the carbon emissions by CH$_4$ and N$_2$O, whose results are listed in column (2)–(3), we can obtain the coefficient of Uncertainty in column (2)–(3) as 0.103, 0.025, respectively,
which passes the significance test at the 5% level, suggesting that EPU exerts a significantly positive influence on carbon emissions.

Table 5. Robustness test—changing the measurement of dependent variables.

|                | (1) GHG  | (2) CH₄  | (3) N₂O |
|----------------|----------|----------|----------|
| L.dependent    | 0.440 ***| 0.416 ***| 0.345 ***|
|                | (9.354)  | (11.024) | (6.126)  |
| Uncertainty    | 0.293 ** | 0.103 ** | 0.025 ** |
|                | (1.978)  | (2.573)  | (2.456)  |
| GDP            | 0.375 ***| 0.070 ** | 0.042    |
|                | (4.960)  | (1.973)  | (0.570)  |
| POP            | 8.176 ***| 1.307    | 5.663 ** |
|                | (3.469)  | (1.269)  | (2.249)  |
| GI             | −0.002   | −0.001   | −0.003   |
|                | (−0.301) | (−0.909) | (−1.251) |
| Density        | −7.379 ***| −0.678   | −5.199 **|
|                | (−3.038) | (−0.649) | (−2.092) |
| Aging          | 0.200    | 0.189 ** | 0.106    |
|                | (0.781)  | (2.112)  | (0.363)  |
| IND            | 0.053 *  | 0.000    | 0.011    |
|                | (1.785)  | (0.019)  | (0.712)  |
| Urban          | −0.379 ***| 0.049    | 0.134 ** |
|                | (−5.771) | (1.333)  | (2.530)  |
| Trade          | 0.013    | 0.007    | 0.030 ***|
|                | (0.670)  | (1.157)  | (3.281)  |
| IFDI           | −0.002   | −0.001   | −0.003   |
|                | (−0.622) | (−0.687) | (−1.202) |
| Year FE        | yes      | yes      | yes      |
| N              | 3021     | 3276     | 3276     |
| AR (1)         | −2.643   | −2.664   | −3.275   |
| AR (1)-P       | 0.008    | 0.008    | 0.001    |
| AR (2)         | 1.231    | 2.418    | 1.630    |
| AR (2)-P       | 0.218    | 0.016    | 0.103    |
| Hansen-P       | 0.403    | 0.388    | 0.323    |

Notes: ***, ** and * indicate statistical significance at the 1%, 5%, and 10% levels, respectively. T-statistics are in parenthesis.

3.2.3. Panel Fixed Effect (FE) Estimation

Thirdly, to avoid the potential bias caused by the estimation method, we re-estimated the model by employing panel FE estimation which incorporates the individual fixed effect and time fixed effect, similar to Chu and Le (2021) [53], who also utilized this estimation to examine the impact of EPU on carbon emissions. We offer the results of panel FE estimation in Table 6. While we control the time fixed effect, individual fixed effect and other variables which may affect the environmental performance, the coefficient of uncertainty in column (1) is 8.246, which is significantly positive, again indicating that the EPU would bring about more carbon emissions. Furthermore, while we capture the EPU by employing the other three variables, whose results are listed in column (2)–(4), we can obtain that the coefficient of uncertainty_trade, uncertainty_absolute, and uncertainty_season in columns (2)–(4) is 0.021, 0.003, and 0.004, respectively, which passes the significance test at the 5% level, also suggesting that EPU exerts a significantly positive effect on carbon emissions. In summary, the results of panel FE estimation are similar to those of GMM estimation in Tables 3 and 4, confirming the reliability of our baseline finding.

3.2.4. New Sample

Thirdly, to avoid the potential bias caused by outliers, we re-constructed new subsamples by dropping observations with values smaller than 10% CO₂ or larger than 10% CO₂, according to Wang et al., (2021a) [42]. We hence re-estimate the GMM based on such
a new sample and display the results in Table 7. It is easily to see that the coefficient of uncertainty, uncertainty_trade, uncertainty_absolute and uncertainty_season are negative and pass the significance test at the 10% level, indicating that the results based on the middle 80% sample also support the positive influence of EPU on carbon emissions.

Table 6. Robustness test—changing the estimation.

|       | (1)       | (2)       | (3)       | (4)       |
|-------|-----------|-----------|-----------|-----------|
| Uncertainty | 8.246 *** | (6.003)   |           |           |
| Uncertainty_trade | 0.021 ** | (2.437)   |           |           |
| Uncertainty_absolute | 0.003 ** | (2.320)   |           |           |
| Uncertainty_season | 0.004 *** | (3.081)   |           |           |

|        | (1)       | (2)       | (3)       | (4)       |
|--------|-----------|-----------|-----------|-----------|
| GDP    | 0.636 *** | (6.743)   | 0.618 *** | (6.724)   |
| POP    | 0.727     | 2.005     | 0.726     | 0.728     |
| GI     | 0.023     | 0.031 **  | 0.023     | 0.023     |
| Density| 0.904     | (1.292)   | 0.906     | 0.904     |
| Aging  | 0.686 *** | (4.390)   | 0.297 *   | 0.686 *** |
| IND    | 0.072     | 0.093 **  | 0.071     | 0.072     |
| Urban  | 0.132     | 1.029 *** | 0.133     | 0.132     |
| Trade  | 0.082     | 0.008     | 0.082 *   | 0.082     |
| IFDI   | −0.010    | 0.018     | −0.010    | −0.010    |
| Year FE| yes       | yes       | yes       | yes       |
| Country FE | yes    | yes       | yes       | yes       |
| Cons   | −19.363 **| −36.814 **| −19.341 **| −19.371 **|

Notes: ***, ** and * indicate statistical significance at the 1%, 5%, and 10% levels, respectively. T-statistics are in parenthesis.

Table 7. Robustness test—middle 80% sample.

|       | (1)       | (2)       | (3)       | (4)       |
|-------|-----------|-----------|-----------|-----------|
| L.CO₂ | 0.433 *** | (7.721)   | 0.522 *** | (14.645)  |
| Uncertainty | 0.026 *** | (3.556)   |           |           |
| Uncertainty_trade | 0.052 *** | (5.262)   |           |           |
| Uncertainty_absolute | 0.005 *** | (2.979)   |           |           |
| Uncertainty_season |          |           |           | 0.086 *   |
| GDP   | 0.743 *** | (7.941)   | 0.329 *** | (12.833)  |
|       |           |           | 0.236 *** | (5.075)   |
|       |           |           | 0.233 *** | (4.960)   |
4. Conclusions

To investigate the influence of EPU on CO$_2$ emissions, we collected the cross-country data for 137 countries from 1970 to 2018 to carry out a GMM system estimation. The baseline result confirms that the EPU would bring about more CO$_2$ emissions, which is also credible when we utilize the new measurement of EPU or the new measurement for CO$_2$ emissions. In addition, while we examine the moderating effect of other factors in EPU’s influence on CO$_2$ emissions, we see that the effect of EPU on air pollution among OECD countries is lower than that in non-OECD countries, suggesting that a higher level of economic development would reduce the environmental adverse effect of EPU. Besides, higher globalization and more international trade would weaken the effect of EPU on CO$_2$ emissions, and better governance would reduce EPU’s influence on CO$_2$ emissions. Aside from this, we also test whether the role of EPU in CO$_2$ emissions varies among different political regimes; the results support that the influence of EPU on globalization is stronger in autocracies than that in democracies, whereas left-wing governments can somewhat reduce EPU’s influence on CO$_2$ emissions.

According to our empirical findings, we offer the following policy implications to improve environmental performance. Firstly, given that EPU would have some negative influence on environmental performance, governments could put in more effort to protect the stability and predictability of economic policies, especially environmentally friendly policies; if they prefer to conduct new policies, the transition should be smooth, especially during election periods or when experiencing significant external shock such as coronavirus; the stability of economic policies can not only spur economic incentives, but also improve environmental performance, which are both essential for national sustainable development. Secondly, among the specific dimension of EPU, trade uncertainty also negatively affects environmental performance, and during the era of post coronavirus, trade uncertainty also increases. To reduce the negative effect of trade uncertainty, policymakers should continue their earlier preference for international trade or trade among domestic...
country, and conduct long-term policies to maintain sustainable environments for trade activities, which are beneficial for environmental protection.

Even if our study comprehensively investigated the impact of economic policy uncertainty on carbon emissions by using 137 countries and reported that EPU causes more carbon emissions, this paper did not uncover the transmission channels through which EPU can affect carbon emissions, nor whether the EPU effect on carbon emissions is constant among different countries, which need to be further investigated in future studies.

**Author Contributions:** H.-J.W. proposed theoretical analysis. Y.G. set up the framework and revised the manuscript. X.-Q.X. performed the calculations and analyzed the data. All of the authors have contributed to the writing and finalizing of the paper. Q.-J.W. putted forward this investigating issue. All authors have read and agreed to the published version of the manuscript.

**Funding:** His research is supported by the China Postdoctoral Science Foundation (funding number: 2021M702961); National Natural Science Foundation of China (funding number: 72088101); National Natural Science Foundation of China (funding number: 71810107001); Humanities and Social Sciences Foundation of Ministry of Education of China (funding number: 21YJC630140); University Science and Technology Innovation Program of He’nan (funding number: 2021-CX-007).

**Institutional Review Board Statement:** Not applicable.

**Informed Consent Statement:** Not applicable.

**Data Availability Statement:** The data used to support the findings of this study are available from the corresponding author upon request.

**Conflicts of Interest:** The authors declare no conflict of interest.

**References**

1. Jordà, Ò.; Singh, S.R.; Taylor, A.M. Longer-run economic consequences of pandemics? *Rev. Econ. Stat.* 2020, 1–29. [CrossRef]
2. Bakas, D.; Triantafyllou, A. Commodity price volatility and the economic uncertainty of pandemics. *Econ. Lett.* 2020, 193, 109283. [CrossRef]
3. Baker, S.R.; Farrokhnia, R.A.; Meyer, S.; Pagel, M.; Yannelis, C. How does household spending respond to an epidemic? Consumption during the 2020 COVID-19 pandemic. *Rev. Asset Pricing Stud.* 2020, 10, 834–862. [CrossRef]
4. Balcilar, M.; Gupta, R.; Segnon, M. The role of economic policy uncertainty in predicting US recessions: A mixed-frequency Markov-switching vector autoregressive approach. *Economics* 2016, 10, 20160027. [CrossRef]
5. Dugiannakis, S.; Filis, G.; Panagiotakopoulos, S. Oil price shocks and uncertainty: How stable is their relationship over time? *Econ. Model.* 2018, 72, 42–53. [CrossRef]
6. Hailemariam, A.; Smyth, R.; Zhang, X. Oil prices and economic policy uncertainty: Evidence from a nonparametric panel data model. *Energy Econ.* 2019, 83, 40–51. [CrossRef]
7. Phan, D.H.B.; Iyke, B.N.; Sharma, S.S.; Affandi, Y. Economic policy uncertainty and financial stability—Is there a relation? *Econ. Model.* 2021, 94, 1018–1029. [CrossRef]
8. Guidolin, M.; La Ferrara, E. The economic effects of violent conflict: Evidence from asset market reactions. *J. Peace Res.* 2010, 47, 671–684. [CrossRef]
9. Kang, W.; Ratti, R.A.; Vespignani, J.L. Oil price shocks and policy uncertainty: New evidence on the effects of US and non-US oil production. *Energy Econ.* 2017, 66, 536–546. [CrossRef]
10. Olanipekun, I.; Olasthinde-Williams, G.; Saint Akadiri, S. Gasoline prices and economic policy uncertainty: What causes what, and why does it matter? Evidence from 18 selected countries. *Environ. Sci. Pollut. Res.* 2019, 26, 15187–15193. [CrossRef] [PubMed]
11. Akron, S.; Demir, E.; Díez-Esteban, J.M.; García-Gómez, C.D. Economic policy uncertainty and corporate investment: Evidence from the US hospitality industry. *Tour. Manag.* 2020, 77, 104019. [CrossRef]
12. Salahuddin, M.; Alam, K.; Ozturk, I.; Sohag, K. The effects of electricity consumption, economic growth, financial development and foreign direct investment on CO2 emissions in Kuwait. *Renew. Sustain. Energy Rev.* 2018, 81, 2002–2010. [CrossRef]
13. Shahbaz, M.; Mahalik, M.K.; Shahzad, S.J.H.; Hammoudeh, S. Does the environmental Kuznets curve exist between globalization and energy consumption? Global evidence from the cross-correlation method. *Int. J. Financ. Econ.* 2019, 24, 540–557. [CrossRef]
14. Kompas, T.; Pham, V.H.; Che, T.N. The effects of climate change on GDP by country and the global economic gains from complying with the Paris climate accord. *Earth’s Future* 2018, 6, 1153–1173. [CrossRef]
15. Coskuner, C.; Paske, M.K.; Olasehinde-Williams, G.; Akadiri, S.S. Economic and social determinants of carbon emissions: Evidence from organization of petroleum exporting countries. *J. Public Aff.* 2020, 20, e2092. [CrossRef]
16. Hambría, W.L.; Saarinen, J.; Moses, O. Climate change policy in a world of uncertainty: Changing environment, knowledge, and tourism in Botswana. *Afr. Geogr. Rev.* 2020, 39, 252–266. [CrossRef]
17. Huo, T.; Li, X.; Cai, W.; Zuo, J.; Jia, F.; Wei, H. Exploring the impact of urbanization on urban building carbon emissions in China: Evidence from a provincial panel data model. *Sustain. Cities Soc.* 2020, 56, 102068. [CrossRef]

18. Jiang, Y.; Zhou, Z.; Liu, C. Does economic policy uncertainty matter for carbon emission? Evidence from US sector level data. *Environ. Sci. Pollut. Res.* 2019, 26, 24380–24394. [CrossRef]

19. Yu, J.; Shi, X.; Guo, D.; Yang, L. Economic policy uncertainty (EPU) and firm carbon emissions: Evidence using a China provincial EPU index. *Energy Econ.* 2021, 94, 105071. [CrossRef]

20. Adams, S.; Adedoyin, F.; Olaniran, E.; Bekun, F.V. Energy consumption, economic policy uncertainty and carbon emissions: causality evidence from resource rich economies. *Econ. Anal. Policy* 2020, 68, 179–190. [CrossRef]

21. Gill, A.R.; Hassan, S.; Viswanathan, K.K. Is democracy enough to get early turn of the environmental Kuznets curve in ASEAN countries? *Energy Environ.* 2019, 30, 1491–1505. [CrossRef]

22. Shabir, M.; Ali, M.; Hashmi, S.H.; Bakhsh, S. Heterogeneous effects of economic policy uncertainty and foreign direct investment on environmental quality: Cross-country evidence. *Environ. Sci. Pollut. Res.* 2021, 1–16. [CrossRef]

23. Adedoyin, F.F.; Zakari, A. Energy consumption, economic expansion, and CO₂ emission in the UK: The role of economic policy uncertainty. *Sci. Total Environ.* 2020, 738, 140014. [CrossRef] [PubMed]

24. Chen, X.; Zhang, S.; Ruan, S. Polycrystalline structure and carbon dioxide emissions: Empirical analysis from provincial data in China. *J. Clean. Prod.* 2021, 278, 123411. [CrossRef]

25. Doğan, E.; Güler, C. How Economic Policy Uncertainty Affect Carbon Emissions: A Case of G-7 Countries. *J. Econ. Bus. Issues* 2021, 1, 33–38.

26. Abbasi, K.R.; Adedoyin, F.F. Do energy use and economic policy uncertainty affect CO₂ emissions in China? Empirical evidence from the dynamic ARDL simulation approach. *Environ. Sci. Pollut. Res.* 2021, 28, 23323–23335. [CrossRef]

27. Liu, Y.; Zhang, Z. How does economic policy uncertainty affect CO₂ emissions? A regional analysis in China. *Environ. Sci. Pollut. Res.* 2021, 1–15. [CrossRef] [PubMed]

28. Baker, S.R.; Bloom, N.; Davis, S.J. Measuring economic policy uncertainty. *Q. J. Econ.* 2016, 131, 1593–1636. [CrossRef]

29. Atsu, F.; Adams, S. Energy consumption, finance, and climate change: Does policy uncertainty matter? *Econ. Anal. Policy* 2021, 70, 490–501. [CrossRef]

30. Pirgaip, B.; Dinçergök, B. Economic policy uncertainty, energy consumption and carbon emissions in G7 countries: Evidence from a panel Granger causality analysis. *Environ. Sci. Pollut. Res.* 2020, 27, 30050–30066. [CrossRef]

31. Shohail, M.T.; Xiuuyuan, Y.; Usman, A.; Majeed, M.T.; Ullah, S. Renewable energy and non-renewable energy consumption: Assessing the asymmetric role of monetary policy uncertainty in energy consumption. *Environ. Sci. Pollut. Res.* 2021, 28, 31575–31584. [CrossRef]

32. Amin, A.; Doğan, E. The role of economic policy uncertainty in the energy-environment nexus for China: Evidence from the novel dynamic simulations method. *J. Environ. Manag.* 2021, 292, 112865. [CrossRef]

33. Anser, M.K.; Apergis, N.; Syed, Q.R. Impact of economic policy uncertainty on CO₂ emissions: Evidence from top ten carbon emitter countries. *Environ. Sci. Pollut. Res.* 2021, 28, 29369–29378. [CrossRef] [PubMed]

34. Ahir, H.; Bloom, N.; Furer, D. The World Uncertainty Index. Available at SSRN 3275033. 2018. Available online: https://papers.ssrn.com/sol3/papers.cfm?abstract_id=3275033 (accessed on 10 October 2021).

35. Shao, S.; Yang, L.; Yu, M.; Yu, M. Estimation, characteristics, and determinants of energy-related industrial CO₂ emissions in Shanghai (China), 1994–2009. *Energy Policy* 2013, 39, 6476–6494. [CrossRef]

36. Qi, X.; Wang, X.; Jin, X.; Wang, Z.M.; Zhang, B.; Wen, C. Will Policy Uncertainty Deteriorate Haze Pollution? A Spatial Spillover Perspective. *Int. J. Environ. Res. Public Health* 2021, 18, 10229. [CrossRef]

37. Wen, J.; Hao, Y.; Feng, G.-F.; Chang, C.-P. Does government ideology influence environmental performance? Evidence based on a new dataset. *Econ. Syst.* 2016, 40, 232–246. [CrossRef]

38. Wang, Q.-J.; Geng, Y.; Xia, X.-Q. Revisited Globalization’s Impact on Total Environment: Evidence Based on Overall Environmental Performance Index. *Int. J. Environ. Res. Public Health* 2021, 18, 11419. [CrossRef]

39. Yang, Q.-C.; Feng, G.-F.; Chang, C.-P.; Wang, Q.-J. Environmental protection and performance: A bi-directional assessment. *Sci. Total Environ.* 2021, 774, 145747. [CrossRef]

40. Verma, A.; Schmidt-Vogt, D.; De Alban, J.D.T.; Lim, C.L.; Webb, E.L. Drivers and mechanisms of forest change in the Himalayas. *Glob. Environ. Chang.* 2021, 68, 102244. [CrossRef]

41. Bottero, A.; D’Amato, A.W.; Palik, B.J.; Bradford, J.B.; Fraver, S.; Battaglia, M.A.; Asherin, L.A. Density-dependent vulnerability of forest ecosystems to drought. *J. Appl. Ecol.* 2017, 54, 1605–1614. [CrossRef]

42. Wang, Q.-J.; Feng, G.-F.; Wang, H.-J.; Chang, C.-P. The impacts of democracy on innovation: Revisited evidence. *Technovation* 2021, 108, 102333. [CrossRef]

43. Usman, M.; Anwar, S.; Yaseen, M.R.; Makhdom, M.S.A.; Kousar, R.; Jahanger, A. Unveiling the dynamic relationship between agriculture value addition, energy utilization, tourism and environmental degradation in South Asia. *J. Public Aff.* 2021, e2712. [CrossRef]

44. Dally, S. Urban bird community composition influenced by size of urban green spaces, presence of native forest, and urbanization. *Urban Ecosyst.* 2018, 21, 1–14. [CrossRef]

45. Li, L.; Liu, J.; Long, H.; de Jong, W.; Youn, Y.-C. Economic globalization, trade and forest transition-the case of nine Asian countries. *For. Policy Econ.* 2017, 76, 7–13. [CrossRef]
46. Mahadevan, R.; Sun, Y. Effects of foreign direct investment on carbon emissions: Evidence from China and its Belt and Road countries. *J. Environ. Manag.* 2020, 276, 111321. [CrossRef]
47. WDI. *World Development Indicators*; World Bank: Washington, DC, USA, 2020.
48. OECD. *OECD Statistics*; World Bank: Washington, DC, USA, 2020.
49. Wang, Q.-J.; Feng, G.-F.; Chen, Y.E.; Wen, J.; Chang, C.-P. The impacts of government ideology on innovation: What are the main implications? *Res. Policy* 2019, 48, 1232–1247. [CrossRef]
50. Ulucak, R.; Khan, S.U.D. Relationship between energy intensity and CO₂ emissions: Does economic policy matter? *Sustain. Dev.* 2020, 28, 1457–1464.
51. Al-Thaqeb, S.A.; Algharabali, B.G. Economic policy uncertainty: A literature review. *J. Econ. Asymmetries* 2019, 20, e00133. [CrossRef]
52. Lin, S.; Sun, J.; Marinova, D.; Zhao, D. Effects of population and land urbanization on China’s environmental impact: Empirical analysis based on the extended STIRPAT model. *Sustainability* 2017, 9, 825. [CrossRef]
53. Chu, L.K.; Le, N.T.M. Environmental quality and the role of economic policy uncertainty, economic complexity, renewable energy, and energy intensity: The case of G7 countries. *Environ. Sci. Pollut. Res.* 2021, 1–17. [CrossRef] [PubMed]