Supplement for ‘Sub-national levels and trends in contraceptive prevalence, unmet need, and demand for family planning in Nigeria with survey uncertainty’

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1. Input Data

In this section we describe the thirteen household surveys that were analyzed for the article ‘Sub-national levels and trends in contraceptive prevalence, unmet need, and demand for family planning in Nigeria with survey uncertainty.’

1.1. Demographic and Health Surveys

We processed the four available Demographic and Health Surveys (DHS) for Nigeria. The 1990 DHS selected 34 households from each of the 299 enumeration areas (EAs), which were stratified by urban/rural status and selected via probability proportional to size (pps), and was designed to produce national and regional (4 regions at that time) estimates [1]. A stratified two-stage sampling design was used for the 2003 DHS to provide estimates of health indicators at the national level, the six health regions, and by urban/rural status. Households were select from a complete listing of households within the 365 selected EAs [2]. Similarly, the 2008 and 2013 DHS employed stratified multi-stage sampling designs consisting of 888 EAs and 904 EAs, respectively [3, 4] and were intended to be representative at the national, regional, and state level. An additional survey was conducted in 1999, however the DHS Program was not centrally involved with the study and does not distribute the collected data.

To calculate unmet need we relied on the recoded definitions used by the United Nations Population Division (UNPD), which applied the revised DHS unmet need definition [5] to previous surveys. To compute the revised unmet need in our analysis, we relied on the UNPD script available at https://github.com/PhilUeff/PDU_FP-Indicators.

To ensure state-level values reflect current geographical divisions, we mapped the GPS coordinates associated with the sampled clusters from each survey onto the 2013 Nigeria DHS shapefile and used their corresponding state names. Each DHS survey contained a few clusters without recorded GPS coordinates (Table 1). For these, we assigned the clusters to the states that they were recorded as belonging to at the time. The assignment is ambiguous for one DHS 1990 cluster because the corresponding recorded state has split between the time of the survey and 2013; this is noted in Table 1.

Table 1 compares national all-women estimates for contraceptive prevalence rate (CPR), modern contraceptive prevalence rate (mCPR), total unmet need (UnN), and demand satisfied by modern methods (DSM), as computed by the DHS survey official reports, the UN recode, and our analysis. While not presented, traditional contraceptive prevalence rate can be computed as CPR - mCPR.

Contraceptive prevalences match among all sources. The differences in unmet need between the DHS official reports and the UN computed values prior to 2012 reflect the revised unmet need calculation. Dashes indicate estimates that were not reported by the DHS.

Table 1: Surveyed DHS clusters without GPS coordinates

| DHS year | Cluster ID | Recorded state | Possible states       |
|----------|------------|----------------|-----------------------|
| 1990     | 1402       | Kwara          | Kwara, Kogi           |
|          | 1554       | Lagos          | Lagos                 |
| 2003     | 170        | Katsina        | Katsina               |
|          | 258        | Lagos          | Lagos                 |
| 2008     | 773        | Rivers         | Rivers                |
|          | 302        | Yobe           | Yobe                  |
|          | 373        | Kano           | Kano                  |
|          | 422        | Kebbi          | Kebbi                 |
| 2013     | 514        | Anambra        | Anambra               |
|          | 557        | Enugu          | Enugu                 |
|          | 569        | Enugu          | Enugu                 |
|          | 639        | Bayelsa        | Bayelsa               |

1.2. Multiple Indicator Cluster Surveys

We used three Multiple Indicator Cluster Surveys (MICS) surveys in our analysis. Thirty EAs were selected within each state for the 2007 MICS, which was designed to provide state and national level estimates of health indicators [6]. A two-stage sampling design was used for the 2011 MICS, which was designed to provide estimates and the state and national level.
Forty EAs were selected with equal probability within each state and all 1,480 EAs were included in field work [7]. The 2016-17 MICS was designed to provide estimates at the national and state level and relied on a two-stage sampling design. Sixty EAs were selected from each state, except for Lagos and Kano which were over sampled at 120 EAs. Of the 2,340 sampled EAs only 2,239 were included in field work because 101 EAs from Borno, Yobe, and Adamawa states were excluded due to security concerns [8].

The UNPD has recoded the 2007 and 2011 MICS surveys, but at the time of this writing, had not yet recoded the 2016 survey. We used the UNPD recode scripts from [https://github.com/PhilUeff/PDU_FP-Indicators](https://github.com/PhilUeff/PDU_FP-Indicators) to process the survey. We used the UN code on the data, we found a small difference in the 2011 estimated unmet need from the value reported by the UNPD. For the 2016 MICS survey, we checked that the variables were encoded in the same form as the 2011 MICS survey. The unavailable estimates from the MICS official reports and 2016 UNPD recode are indicated with dashes. Finally, certain questions needed to calculate unmet need were not asked in the 2007 MICS survey. As a result, unmet need and demand satisfied for this survey are marked with ‘NA’. The MICS official reports only present the estimates over currently-married/in-union women. Table 4 displays the estimates reported by MICS, calculated by the UN, and calculated via our direct analysis of the microdata.

### Table 2: DHS surveys: comparing DHS reported figures, UN computed figures, and our computed figures at the national level for all women. All values are reported in (%).

| DHS year | CPR  | mCPR | UnN  | DSM  | CPR  | mCPR | UnN  | DSM  | CPR  | mCPR | UnN  | DSM  |
|----------|------|------|------|------|------|------|------|------|------|------|------|------|
| 2013     | 16.0 | 11.1 | 12.7 | 38.8 | 16.0 | 11.1 | 12.7 | 38.5 | 16.0 | 11.1 | 12.7 | 38.8 |
| 2008     | 15.4 | 10.5 | 15.7 | –    | 15.4 | 10.5 | 16.1 | 33.2 | 15.4 | 10.5 | 16.1 | 33.2 |
| 2003     | 13.3 | 8.9  | 13.6 | –    | 13.3 | 8.9  | 14.6 | 32.0 | 13.3 | 8.9  | 14.6 | 32.0 |
| 1990     | 7.5  | 3.8  | –    | –    | 7.5  | 3.8  | 17.1 | 15.2 | 7.5  | 3.8  | 17.1 | 15.2 |

1.4. National Nutrition and Health Surveys

We analyzed the microdata from two National Nutrition and Health Surveys (NNHS) from 2014 and 2015. Both surveys were conducted using the Standardized Monitoring and Assessment of Relief and Transition (SMART) methods and relied on two-stage cluster and designed to provide estimates at the national and state level [15, 16]. Unfortunately, the NNHS does not include the variables required to calculate unmet need nor parity, so the NNHS data was only included in the estimates of mCPR for all women. Additionally, the survey reports highlight that in 9 Local Governmental Areas (LGAs) were excluded for security reasons and thus results for Borno are not representative of the whole state.

2. Modeling family planning indicators

Family planning indicators in Nigeria are estimated from multiple complex surveys, sometimes in the same year and can be quite noisy between surveys at the state level. In an effort to understand the underlying population rates from which these survey observations were drawn, we are adapting a previously developed two-step space-time smoothing approach that acknowledges complex sampling designs to model the family planning indicators.

The first step of our approach requires estimating the state (i), year (t), and survey (s) specific direct estimates of proportions and corresponding variance via the Hajek estimator [17]

\[
\hat{p}_{its} = \sum_{j} x_{jits} w_{jits} / \sum_{j} w_{jits}
\]

where \( x_{jits} \) is the binary indicator for modern contraception, traditional contraception, unmet need, or demand satisfied for woman \( j \), in sampled in area \( i \), time \( t \), and survey \( s \), and \( w_{jits} \) is her corresponding sampling weight.

In the second step we use a three-stage Bayesian hierarchical model. In the first stage we rely on the working likelihood based on the asymptotic distribution,

\[
Y_{its} \sim N(\eta_{its}, \tilde{V}_{its,DES})
\]
Table 3: MICS surveys: comparing MICS reported figures, UN computed figures, and our computed figures at the national level for all women. All values are reported in (%).

| MICS year | CPR | MCPR | UnN | DSM | CPR | MCPR | UnN | DSM | CPR | MCPR | UnN | DSM |
|-----------|-----|------|-----|-----|-----|------|-----|-----|-----|------|-----|-----|
| 2016      |     |      |     |     |     |      |     |     | 11.8| 9.5  | 26.8| 24.6|
| 2011      |     |      |     |     | 18.3| 12.5 | 20.2| 32.6|     |     | 18.3| 12.5 | 21.0| 31.9|
| 2007      |     |      | NA  | NA  | 14.6| 11.1 | NA  | NA  |     |     | 14.6| 11.1 | NA  | NA  |

Table 4: PMA surveys: comparing PMA2020 reported figures and our computed figures at the national level for all women. All values are reported in (%).

| PMA year | State   | CPR | MCPR | UnN | DS | CPR | MCPR | UnN | DS |
|----------|---------|-----|------|-----|----|-----|------|-----|----|
| 2017     | Anambra | 24.2| 13.9 | 12.0| 38.3|     | 24.1| 13.8| 12.0| 38.2|
|          | Kaduna  | 17.4| 15.6 | 22.4| 39.2|     | 17.4| 15.6| 22.4| 39.2|
|          | Kano    | 6.2 | 4.4  | 24.6| 14.4|     | 6.2 | 4.4  | 24.6| 14.4|
|          | Lagos   | 29.7| 20.6 | 15.2| 45.8|     | 29.7| 20.6 | 15.2| 45.8|
|          | Nasarawa| 18.9| 16.9 | 21.0| 42.3|     | 18.9| 16.9 | 21.0| 42.3|
|          | Rivers  | 29.0| 17.7 | 17.2| 38.2|     | 29.0| 17.7 | 17.2| 38.2|
|          | Taraba  | 14.1| 10.5 | 24.4| 27.3|     | 14.1| 10.5 | 24.4| 27.3|
| 2016     | Anambra | 25.1| 15.6 | 14.4| 39.6|     | 25.9| 16.1 | 14.4| 40.4|
|          | Kaduna  | 15.1| 13.9 | 26.2| 33.7|     | 15.4| 14.1 | 26.2| 34.0|
|          | Kano    | 5.6 | 4.8  | 30.3| 13.5|     | 5.6 | 4.9  | 30.2| 13.5|
|          | Lagos   | 26.4| 19.7 | 15.6| 46.8|     | 26.0| 19.7 | 15.6| 46.9|
|          | Nasarawa| 18.9| 16.6 | 18.0| 44.9|     | 18.9| 16.6 | 18.0| 44.9|
|          | Rivers  | 27.5| 19.4 | 16.4| 44.2|     | 28.8| 19.6 | 16.4| 43.7|
|          | Taraba  | 12.9| 9.9  | 27.1| 24.9|     | 12.6| 10.0 | 27.1| 25.0|
| 2015     | Kaduna  | 14.5| 13.2 | 24.9| 33.5|     | 14.7| 13.5 | 24.9| 34.2|
|          | Lagos   | 27.8| 21.0 | 17.8| 46.1|     | 27.8| 21.4 | 17.8| 46.4|
| 2014     | Kaduna  | 8.7 | 8.4  | 28.2| 22.8|     | 8.7 | 8.5  | 28.2| 22.9|
|          | Lagos   | 17.8| 16.5 | 19.3| 44.5|     | 17.3| 16.7 | 19.2| 44.8|

where $Y_{its} = \logit(p_{its})$ and $\hat{V}_{its,DES}$ is the design-based variance of $Y_{its}$. This pseudo-likelihood for a binary variable was initially developed to estimate zipcode-level smoking prevalence in Washington state \[18\] relying on a single survey instrument. This model was then adapted for the complex indicator of under-five child mortality \[19\] and extended the classic inseperable spatio-time models of \[20, 21\] to include survey-specific random effects.

At the second stage we assume

$$
\eta_{its} = \mu + \gamma_t + \alpha_t + \beta_s + \delta_{it} + \phi_{is} + \psi_{ts},
$$

where $\mu$ is a shared intercept, the temporally structured random effects $\alpha$ are assigned second order random walk priors \[22, 21\], the spatial random effects $\beta$ are assigned either the Besag, York, and Mollie (BYM) \[23\] or the scaled BYM \[24\], the space-time interactions $\delta$ are assigned the ‘type II’ temporally structured interaction \[21\], and the independent random effects
\(\gamma, \nu, \phi, \) and \(\psi\) are assigned independent mean-zero Normal distributions. Note, if \(\theta \sim BYM\) then \(\theta_i = U_i + V_i\), where \(U\) is assigned an intrinsic conditional autoregression prior (ICAR) \(^{[25]}\) and \(V\) are assigned an independent mean-zero Normal prior. At the third stage of the model we assume a default diffuse prior for \(\mu\) and Gamma(1,5e−5) priors were assigned to \(\sigma^{-2}_\gamma, \sigma^{-2}_\nu, \sigma^{-2}_\phi, \) and \(\sigma^{-2}_\psi\).

Finally, to generate our estimates of the underlying rates we adopt the approach of Mercer (2015) and draw \(k=1,...1000\) posterior samples for all parameters and calculate
\[ \eta_t^{(k)} = \mu + \gamma_i^{(k)} + \alpha_t^{(k)} + \theta_i + \phi_t. \]

The posterior median is used as the estimate and the uncertainty intervals are defined by the 2.5% and 97.5% quantiles.

3. Model Selection

The twelve possible models considered for mCPR, traditional contraceptive prevalence, unmet need, and demand satisfied are described in Table 6. Each model was fit for mCPR, traditional contraceptive prevalence, unmet need, and demand satisfied for all women and the four age and parity combinations. For each model fit we calculated the sum of the log conditional predictive ordinate (LCPO) \(^{[26]}\), the deviance information criteria (DIC) \(^{[27]}\), and the Watanabe-Akaike information criterion (WAIC) \(^{[28]}\). The results for mCPR and unmet need for all women are shown in Table 6 and Table 7, respectively with with the lowest DIC and WAIC and highest LCPO in bold. Table 5 shows the selected model for all outcomes and age-parity groups.

Table 5: Random effects models considered for time \(t\), state \(i\), and survey \(s\). All models contain an intercept \(\mu\), temporally independent \(\gamma_i \sim N(0, \sigma^2_\gamma)\), and temporally structured effect \(\phi_t \sim RW2(\sigma^2_\phi)\). The survey and survey interaction random effects were assigned mean-zero Normal priors. In models designated with ‘a’ \(\theta_i \sim BYM(\sigma^2_\theta)\) and in models designated ‘b’ \(\theta_i \sim BYM2(\sigma^2_\theta)\).

| Model | outcome |
|-------|---------|
| 1     | \(\mu + \alpha_t + \gamma_i + \theta_i\) |
| 2     | \(\mu + \alpha_t + \gamma_i + \theta_i + \delta_{it}\) |
| 3     | \(\mu + \alpha_t + \gamma_i + \theta_i + \delta_{it} + \nu_t\) |
| 4     | \(\mu + \alpha_t + \gamma_i + \theta_i + \delta_{it} + \nu_t + \phi_{it}\) |
| 5     | \(\mu + \alpha_t + \gamma_i + \theta_i + \delta_{it} + \nu_t + \psi_{it}\) |
| 6     | \(\mu + \alpha_t + \gamma_i + \theta_i + \delta_{it} + \nu_t + \phi_{it} + \psi_{it}\) |

4. Decomposition of Variance

To assess the relative importance of the individual random effects in explaining the variance in the observed data we calculated the percent of the total variance described by each random effect. Table 5 shows the percent of the variance described by each random effect where the total variance for model 6 is
\[ \sigma^2_\theta + \sigma^2_\gamma + \sigma^2_\nu + \sigma^2_\phi + \sigma^2_\psi + \sigma^2_\phi \]
and for outcomes and age-parity groups using model 4, the \(\sigma^2_\phi\) term is removed.
Table 8: Selected models.

| Age-parity group | outcome            | selected model |
|------------------|--------------------|----------------|
| All women        | mCPR               | 6b             |
|                  | Unmet Need         | 6a             |
|                  | tCPR               | 4b             |
|                  | Demand Satisfied   | 6b             |
| 15-24yo, 0       | mCPR               | 6a             |
|                  | Unmet Need         | 4a             |
|                  | tCPR               | 6a             |
|                  | Demand Satisfied   | 6a             |
| 15-24yo, 1+      | mCPR               | 6b             |
|                  | Unmet Need         | 4b             |
|                  | tCPR               | 6b             |
|                  | Demand Satisfied   | 6a             |
| 25-49yo, 0       | mCPR               | 6b             |
|                  | Unmet Need         | 6b             |
|                  | tCPR               | 6b             |
|                  | Demand Satisfied   | 6a             |
| 25-49yo, 1+      | mCPR               | 6b             |
|                  | Unmet Need         | 6b             |
|                  | tCPR               | 4b             |
|                  | Demand Satisfied   | 4b             |

5. Trends in family planning indicators in PMA2020 states

The observed data and smoothed estimates for mCPR, traditional contraceptive prevalence, unmet need, and demand satisfied in the seven states (Anambra, Kaduna, Kano, Lagos, Nasarawa, Rivers, and Taraba) are shown in Figures 1 through 7.

6. Maps of family planning indicators for all age-parity subgroups

Figures 8, 9, 10, and 11 show the 2017 state-level estimates or the four age-parity subgroups for mCPR, unmet need, traditional contraceptive prevalence, and demand satisfied, respectively.

7. Small area estimation at the second administrative level

The health policies are often enacted at the second administrative level, called local government areas (LGAs) in Nigeria, and ideally we would provide estimates at the LGA-level. Unfortunately, MICS, DHS, PMA2020, and NNHS do not include the LGA locations for the sampled clusters. However, the DHS have provided jittered GPS locations which can be assigned LGA membership. In this section we investigate and compare the precision attained at the state- and LGA-level. We apply the methods described in Section 5, removing the survey-related random effects, to the DHS data from 1990, 2003, 2008, and 2013.

The left panel of Figure 12 display the state-level estimates and 95% credible intervals for mCPR and the right panel displays the same, but at the LGA-level. From the figure it is clear that in many of the LGAs the uncertainty is intolerably large. This is not a surprising result, as many LGAs have few or no sampled clusters. By contrast differences between states and within states over time can be reliably estimated; see Figure 12 in the Appendix and Figure 4 in the main article, respectively. With currently available data it is not possible to generate sub-state level estimates for FP indicators with reasonable precision. FP estimates at the LGA rely exclusively on DHS data and are thus not timely nor precise.

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Table 9: Percent of total variance described by each random effect, where the interpretation is $\sigma_\phi^2$ for spatial, $\sigma_\gamma^2$ for independent temporal, $\sigma_\delta^2$ for structured (RW2) temporal, $\sigma_\psi^2$ for space-time interaction, $\sigma_\alpha^2$ for surveys, $\sigma_\alpha^2$ for survey-space interactions, and $\sigma_\psi^2$ for survey-time interaction.

| Age-parity group | outcome       | $\sigma_\phi^2$ | $\sigma_\gamma^2$ | $\sigma_\delta^2$ | $\sigma_\alpha^2$ | $\sigma_\psi^2$ | $\sigma_\psi^2$ |
|------------------|---------------|-----------------|-------------------|-------------------|-----------------|-----------------|-----------------|
| All women        | mCPR          | 74.5            | <0.05             | <0.05             | 0.5             | <0.05           | 23.8            | 1.2             |
|                  | Unmet Need    | <0.05           | <0.05             | <0.05             | 0.6             | 14.8            | 79.3            | 5.3             |
|                  | tCPR          | 62.2            | <0.05             | 12.5              | <0.05           | 24.7            | -               |
|                  | Demand Satisfied | 43.4        | <0.05             | <0.05             | 0.2             | <0.05           | 55.7            | 0.7             |
| 15-24yo, 0       | mCPR          | 97.5            | <0.05             | <0.05             | <0.05           | <0.05           | 2.0             | 0.5             |
|                  | Unmet Need    | <0.05           | <0.05             | 7.1               | <0.05           | <0.05           | 92.9            | -               |
|                  | tCPR          | <0.05           | <0.05             | <0.05             | 0.1             | <0.05           | 99.9            | <0.05           |
|                  | Demand Satisfied | <0.05       | <0.05             | <0.05             | <0.05           | <0.05           | 93.0            | 7.0             |
| 15-24yo, 1+      | mCPR          | 79.1            | <0.05             | <0.05             | <0.05           | <0.05           | 20.0            | 0.8             |
|                  | Unmet Need    | 56.5            | <0.05             | 24.3              | <0.05           | <0.05           | 19.2            | -               |
|                  | tCPR          | 80.3            | <0.05             | <0.05             | <0.05           | <0.05           | 19.7            | <0.05           |
|                  | Demand Satisfied | <0.05       | <0.05             | <0.05             | <0.05           | <0.05           | 99.9            | <0.05           |
| 25-49yo, 0       | mCPR          | 69.1            | <0.05             | 2.8               | <0.05           | <0.05           | 28.2            | <0.05           |
|                  | Unmet Need    | <0.05           | <0.05             | <0.05             | <0.05           | <0.05           | 32.7            | 67.2            |
|                  | tCPR          | 99.9            | <0.05             | <0.05             | <0.05           | <0.05           | <0.05           | <0.05           |
|                  | Demand Satisfied | <0.05       | <0.05             | <0.05             | <0.05           | <0.05           | 99.9            | <0.05           |
| 25-49yo, 1+      | mCPR          | 97.7            | <0.05             | <0.05             | 0.1             | <0.05           | 2.1             | 0.1             |
|                  | Unmet Need    | 1.4             | <0.05             | 0.3               | 19.8            | 70.1            | 8.4             |
|                  | tCPR          | 62.1            | <0.05             | 13.6              | 0               | 24.0            | -               |
|                  | Demand Satisfied | 46.2        | <0.05             | 0.1               | 0               | 53.6            | -               |

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Figure 1: Data and smoothed estimates for family planning indicators in Anambra state.
Figure 2: Data and smoothed estimates for family planning indicators in Kaduna state.
Figure 3: Data and smoothed estimates for family planning indicators in Kano state.
Figure 4: Data and smoothed estimates for family planning indicators in Lagos state.
Figure 5: Data and smoothed estimates for family planning indicators in Nasarawa state.
Figure 6: Data and smoothed estimates for family planning indicators in Rivers state.
Figure 7: Data and smoothed estimates for family planning indicators in Taraba state.
Figure 8: Smoothed estimates of mCPR by age and parity group for 2017.
Figure 9: Smoothed estimates of unmet need by age and parity group for 2017.
Figure 10: Smoothed estimates of traditional CPR by age and parity group for 2017.
Figure 11: Smoothed estimates of demand satisfied by age and parity group for 2017.
I. State-level estimates

II. Local government area-level estimates

Figure 12: Maps and ranked posterior medians and 95% credible intervals for 2013 at the I. State- and II. Local government area-level (LGA). To generate the map of Nigeria at the LGA level, we utilized the publicly available shapefiles at [30].