Survival analysis of time to cure on multidrug resistance tuberculosis patients in Amhara region, Ethiopia

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Abstract

Background: Multidrug-resistant tuberculosis (MDR-TB) is caused by bacteria that are resistant to the most effective anti-tuberculosis drug. The MDR-TB is an increasing global problem and the spread of MDR-TB has different recovery time for different patients. Therefore, this study aimed to investigate the recovery time of MDR-TB patients in Amhara region, Ethiopia.

Method: A retrospective study was carried out in seven hospitals having MDR-TB treatment center of Amhara region, Ethiopia from September 2015 to February 2018. An accelerated failure time and parametric shared frailty models were employed.

Results: The study revealed that the recovery time of MDR-TB patients in Amhara region was 21 months. Out of the total MDR-TB patients, 110 (35.4%) censored and 201 (64.6%) cured of MDR-TB. The clustering effect of frailty model was hospitals and the Weibull-gamma shared frailty model was selected among all and hence used for this study. The study showed that extra pulmonary MDR-TB patients had longer recovery time than that of smearier pulmonary MDR-TB patients in Amhara region, Ethiopia. According to this study, male MDR-TB patients, MDR-TB patients with co-morbidity and clinical complication were experiencing longer recovery time than that of the counter groups. This study also showed that MDR-TB patients with poor adherence had longer recovery time than those with good adherence MDR-TB patients.

Conclusion: Among different factors considered in this study, MDR-TB type, clinical complication, adherence, co-morbidities, sex, and smoking status had a significant effect on recovery time of MDR-TB patients in Amhara region, Ethiopia.

In conclusion, the Regional and Federal Government of Ethiopia should take immediate steps to address causes of recovery time of MDR-TB patients in Amhara region through encouraging adherence, early case detection, and proper handling of drug-susceptibility according to WHO guidelines.

Keywords: Amhara region, Multidrug-resistance tuberculosis, Average recovery time, Accelerated failure time models, Parametric shared frailty

Background

Tuberculosis (TB) is the ninth leading cause of death worldwide and the leading cause of a single infectious agent, ranking above HIV/AIDS [1]. An estimated 10.4 million people had TB in 2016: 74% were in Africa and 56% were in India, Indonesia, China, the Philippines and Pakistan [1]. Globally in 2016, 6.6 million people with tuberculosis (TB) of these, just over 6.3 million had new or relapsed TB [1].

According to WHO multidrug-resistant tuberculosis (MDR-TB) is caused by bacteria that are resistant to the most effective anti-tuberculosis drugs (isoniazid and rifampicin) [1–3]. MDR-TB results either from primary infection or develop in the course of treatment of a patient due to human error, poor supply management, poor quality anti-TB drugs and/or improper treatment [2, 4, 5]. In addition, poor infection control practice has also been identified as a major factor for the spread of MDR-TB and MDR-TB has different recovery time for...
different patients [6]. MDR-TB is being an increasing global problem, and in 2016, 153,119 cases were notified from which 129,689 enrolled for treatment, of which only 22% started treatment [1]. Assefa et al. also noted that 3.7% new and 20% previously treated MDR-TB cases were identified [7].

The burden and incidence of MDR-TB is increasing and varying significantly from country to country. The countries with the largest number of MDR/RR-TB cases (47% of the global total) were China, India and the Russian Federation [1]. The highest (28%) rate of new MDR-TB cases are from the Soviet Union including regions that share borders with the European Union [8]. In Africa, an estimated 69,000 cases emerged of which about 1.2% were new. 12% of re-treatment cases were from Ethiopia of which 1.6 and 12% MDR-TB patients were new and previously treated TB cases respectively [9]. In addition, Ethiopia is one among the 20 countries with the highest absolute estimated number of incidents of TB and MDR-TB [1]. In comparison to drug-susceptible TB, that takes about 6 to 9 months to treat, recommended treatment for MDR-TB lasts 18 to 24 months or longer [10], and requires the second-line medicines that are not effective as first-line medicines commonly prescribed to treat TB [10]. Previous studies indicated that drug-resistant strains of *Mycobacterium tuberculosis* are of great concern as they are more toxic and more expensive than the first-line regimen [11]. Hence, monitoring closely patients while they take these drugs is critical, as the medications may lead to other serious health problems such as damage to the kidneys, liver, or heart; loss of vision or hearing; and changes in behavior or mood including depression or psychosis [12].

As Ethiopia is one of the 20 high burdens MDR-TB countries and MDR-TB has been a major health problem of the society in the Amhara region, a strategy to provide culture and drug susceptibility testing services has been designed [2, 7]. Even though various studies on the prevention and control of the cross-transmission of healthcare-acquired infections between hospitalized patients have been carried out, the prevalence is still increasing [13, 14]. Importantly, the appearance and transmission of MDR-TB is increasing in hospitals worldwide [15]. MDR-TB poses therapeutic difficulties in the twenty-first century, with only a few antibiotics continuing effective [16]. Consequently, controlling and preventing the emergence and overflow of MDR-TB organisms is of vital importance. Thus, the aim of this study is to investigate the recovery time of MDR-TB patients in Amhara region, Ethiopia using accelerated failure time and parametric shared frailty models.

**Method**

**Data source, sampling design, and sample size**

A retrospective study is carried out in seven hospitals of Amhara region which have MDR-TB treatment center from September 2015 to February 2018. Amhara National regional state is the second largest region of Ethiopia with an estimated area of 159,173.66 km² and an estimated population density of 108.2 [17]. The region is divided into 11 administrative Zones, and 151 Woredas (128 rural and 23 urban). Seven of the 17 hospitals (University of Gondar Teaching Referral Hospital, Borumeda Hospital, Debre Markos Referral Hospital, Woledia Hospital, Metema Hospital, Debre Tabor Hospital, and Debre Berhan Referral Hospital) were included in this study whereas the others were excluded because they have no MDR-TB treatment. In addition, patients that have no full history about their epidemiological, clinical and laboratory results were excluded from the study using exclusion criteria. The sample size was determined using at 95% CI with a prevalence of MDR-TB rate of 15% [18] and a margin error of 0.036. Then a total sample of 396 (377 and 5% contingency) MDR-TB patients were considered using simple random sampling methods. Further discussions on sampling are available at Cochran [19]. Thus, a total sample of 311 MDR-TB patients that fulfill the inclusion-exclusion criteria was considered applying simple random sampling methods. Data were collected by trained nurses under the supervision of investigators and the data quality had been checked for their completeness, consistency, and accuracy by investigators every day (Fig. 1).

**Measurements**

The response variable in this study is defined to be the treatment period from the starting time of MDR-TB treatment up to the time of its cure. The event of interest was recovery from MDR-TB (1 = recovery and 0 = censored). The predictor variables that are included in this study were background characteristics of MDR-TB patients and history of epidemiological, clinical and laboratory results (Table 1). Data were entered and cleaned using SPSS-22 and analyzed using STATA-14.

**Statistical analysis**

Survival analysis is the analysis of statistical data in which the outcome variable of interest is time until an event occurs. Survival data are almost always incomplete and called censoring that may be a right censor, left censoring and interval censoring. The most common are right censoring that happens when a subject follow-up times to terminate before the outcome of interest observed [20]. In any applied set, a survival data can summarize through life tables [21], Kaplan-Meier Survival functions [22] and median survival time [23, 24]. Besides estimating the survival functions, comparing two or more estimated survival curves is the most frequently used statistical tool of recent clinical research [25]. The simplest way of comparing the survival times obtained from two or more groups are the Kaplan-Meier curves.
The following flowchart shows the sampling procedure.

- Identified MDR-TM Patients (N=396)
- Incomplete History (N=62)
- Eligible MDR-TB Patients (N=334)
- Exclusion as Age <18 Years (N=17)
- Unable to access (6)

Analyze Sample (n=311)

**Table 1** Predictors considered in the study

| Variables                     | Categories                                      |
|-------------------------------|------------------------------------------------|
| Sex                           | 0 = Male                                        |
|                               | 1 = Female                                      |
| Age                           | Continuous (year)                               |
| Residence                     | 0 = urban,                                      |
|                               | 1 = Rural                                       |
| Marital status                | 0 = Married,                                    |
|                               | 1 = Single,                                     |
|                               | 2 = Divorced,                                   |
|                               | 3 = widowed                                     |
| HIV status                    | 0 = HIV Negative                                |
|                               | 1 = HIV Positive                                |
| Smoking history               | 1 = Smokers,                                    |
|                               | 0 = Non-smokers                                 |
| Adherence based on Pill count | 2 = Good                                        |
|                               | 1 = Fair                                        |
|                               | 0 = Poor                                        |
| Comorbidities                 | 1 = Yes                                         |
|                               | 0 = No                                          |
| Previous drug use history     | 1 = yes                                         |
|                               | 0 = No                                          |
| MDR-TB type                   | 0 = pulmonary                                   |
|                               | 1 = Extra pulmonary                             |
| Baseline weight               | Continuous (kg)                                 |
| Occupation                    | 0 = Employee                                    |
|                               | 1 = Farmer                                      |
|                               | 2 = Merchant                                    |

HIV: human immunodeficiency virus, MDR-TB: multidrug-resistant tuberculosis, kg: kilogram
and log-rank test [26–28]. However, to explore the relationship between the survival experience of individual and explanatory variables, an approach based on statistical modeling has been used [29]. Also used with a modeling approach to the analysis of survival data are the Cox Proportional Hazard [30], Accelerated Failure Time [30] and parametric shared frailty models [31].

Accelerated failure time model

Parametric models are very applicable to analyze survival data; there are relatively few probability distributions of the survival time that can be used with these models. In Accelerated Failure Time (AFT) models, we measure the direct effect of the explanatory variables on the survival time, instead of a hazard. This allows an easier interpretation of the results because the parameters measure the effect of the correspondent covariate on the mean survival time. For AFT models it is common to use the log-linear representation:

\[ Y_i = \log T_i = \mu + \beta_1 x_{i1} + \ldots + \beta_p x_{ip} + \sigma \varepsilon_i, \]

where \( \log T_i \) represents the log-transformed survival time, \( \mu \) is the intercept and \( \sigma \) is the scale parameter, the \( x_{i1}, \ldots, x_{ip} \) are the explanatory variables with the coefficients \( \beta \) reflecting the effect that each explanatory variable have on the survival time and estimated by maximum likelihood method using a Newton-Raphson procedure and \( \varepsilon_i \) is the error term which is assumed to follow a specific distribution such as Weibull [32], log-normal [33], log-logistic [30] and gamma [34] among money.

Parametric shared frailty models

In order to account for unobserved heterogeneity, the frailty term was first introduced by Hougaard in 1991 [35], which is an extension of proportional hazards. In a shared frailty model, lifetimes of a group of observations in the same cluster share the same level of frailty with each other [36] that the common frailty variance measures of dependence among lifetimes within a cluster.

Suppose that there are \( i \) clusters and each cluster \( i \) have \( n_i \) observations where \( \sum_{i=1}^{I} n_i = n \) is the total sample size and \( t_{ij} = \min(c_{ij}, t_{ij}) \) is the observed failure time of a right censoring scheme for \( t^{th} (k = 1, \ldots, n) \) observation in \( i^{th} \) cluster and \( c_{ij} \) is the censoring time, where \( t_{ij} \) and \( c_{ij} \) are independent random variables [31]. Then the observed censoring indicator \( d_{ij} \) is equal to 1 if \( t_{ij} < c_{ij} \) and 0 otherwise and conditional on frailty \( z_i > 0 \) and \( X_{ij} \), the hazard function of \( i^{th} \) cluster has the form: \( h(t_{ij}X_{ij} z_i) = z_i h_0(t_{ij}) \exp(\beta X_{ij}) \), where \( h_0(.) \) is the baseline hazard function, \( X_{ij} \) is a vector of observed predictors for the \( k^{th} \) observation and \( \beta \) is a vector of regression parameters.

The frailties, \( z_i \), are i.i.d. variables with the common probability density function \( g(z_i) \). Various studies were done on the choice of continuous distribution of frailty random variables such as Gamma [34], inverse Gaussian [36], log-normal [33] and positive stable [37] and few studies done on the discrete distributions [38]. However, the Gamma distribution is the most common and widely used in literature for determining the frailty effect, which acts multiplicatively on the baseline hazard [31]. Due to its computational convenience, Gamma distribution of mean 1 and variance \( \theta \) used as the frailty distribution for this study.

The probability density functions of one parameter Gamma distribution is given as:

\[ f(x; \theta) = \frac{1}{\Gamma(\theta)} x^{\theta-1} \exp(-x), \quad \theta > 0 \]

The larger value of \( \theta \) indicates the greater degree of heterogeneity among lifetimes within a cluster. Under the concept of gamma frailty, unconditional survival functions as cluster \( i \) obtained by integrating the conditional survival function of the Gamma distribution.

Once the parametric form of baseline hazard specified, the unconditional likelihood function can be easily derived [31, 39, 40] as:

\[ L(\psi, \theta, \beta) = \prod_{i=1}^{I} \frac{1}{\theta^s(t_{ij}) \psi} \exp(-\theta^s(t_{ij}) \psi \exp(\beta X_{ij})) \exp(\beta X_{ij}) \]

where \( i = 1, \ldots, n \) and \( k = 1, \ldots, n_i \). \( \psi \) is the vector of baseline hazard parameters, \( \Lambda_0(.) = \int_0^{t_{ij}} h_0(s) ds \) is the common cumulative baseline hazard and \( \beta \)'s are the regression parameters that are estimated based on the marginal likelihood in which the frailties have been integrated out by averaging the conditional likelihood with respect to the frailty distribution [31, 40].

Models comparison and diagnostics

There are several methods of model selection. But in this study, AIC [41] criteria were used to compare various candidate models and the model with the smallest AIC value is considered as a better fit [42]. After a model fitted, the adequacy of the fitted model needs to be assessed. The methods that involved the model checking for this study used evaluation of the Parametric Baselines and the Cox-Snell Residuals [43].

Result

Data exploratory

The study revealed that the recovery time of MDR-TB patients in Amhara region was 21 months with minimum and maximum recovery time of 17 and 25 months. Table 2 showed that out of the total MDR-TB patients, 110 (35.4%) were censored and 201 (64.6%) were cured of MDR-TB.
The Kaplan-Meier survival function is an important tool for analyzing censored data [31, 44]. The Kaplan-Meier estimator survival curve depicted the overall estimated survivor function and different groups of predictors. Clearly, the overall estimated survivor function showed that MDR-TB patients cured after the 20 months treatment. In addition, the survival ability of patients was the difference between sex, HIV status, smoking, and clinical complication whereas, between residence, marital status, education level, and occupation did not show a clear difference (Fig. 2a-f).

The observed difference in survival experiences in different patient groups was also assessed using the Long-rank and Breslaw test. Table 3 shows that there is a significant survival ability difference between sex, disease categories, drug use history, clinical complication, HIV status, adherence, smoking status, and comorbidity status at 5% significant level. In addition, sex with HIV and drug use history with the smoker is statistical significant interaction effect on MDR-TB patients.

### Table 2 Characteristics of MDR-TB patients in Amhara region, Ethiopia

| Predictors             | Labels          | Status of Patients |              |              |              |
|------------------------|-----------------|--------------------|--------------|--------------|--------------|
|                        |                 | Total Cured/Event (%) | Censored |             |              |
| Sex                    | Female          | 133                | 79 (39.3)   | 54           |
|                        | Male            | 178                | 122 (60.7)  | 56           |
| Residence              | Urban           | 167                | 106 (52.7)  | 61           |
|                        | Rural           | 144                | 95 (47.3)   | 49           |
| Treatment Center       | Gondar          | 110                | 57 (28.4)   | 53           |
|                        | Boromeda        | 48                 | 33 (16.4)   | 15           |
|                        | Debre Markos    | 46                 | 32 (15.9)   | 14           |
|                        | Woldeia         | 33                 | 27 (33)     | 6            |
|                        | Metema          | 30                 | 20 (10)     | 10           |
|                        | Debre Tabore    | 25                 | 18 (9)      | 7            |
|                        | Debre Berhan    | 19                 | 14 (7)      | 5            |
| Disease category       | Pulmonary       | 262                | 194 (96.5)  | 68           |
|                        | Extra Pulmonary | 49                 | 7 (3.7)     | 42           |
| Drug using history     | Yes             | 233                | 147 (73.1)  | 86           |
|                        | No              | 78                 | 54 (26.9)   | 24           |
| Clinical               | Complication    | 129                | 67 (29.9)   | 62           |
|                        | Not Complication| 182               | 134 (70.1)  | 48           |
| HIV status             | HIV Positive    | 74                 | 42 (20.9)   | 32           |
|                        | HIV Negative    | 237                | 159 (79.1)  | 78           |
| Smoking history        | Yes             | 72                 | 50 (57.1)   | 22           |
|                        | No              | 239                | 151 (24.9)  | 88           |
| Comorbidities          | Yes             | 27                 | 18 (66.7)   | 9            |
|                        | No              | 284                | 183 (64.4)  | 101          |
| Adherence              | Good            | 226                | 176 (77.9)  | 50           |
|                        | Fair            | 39                 | 19 (48.7)   | 20           |
|                        | Poor            | 46                 | 6 (13)      | 40           |
| Occupation             | 0 = Employee    | 34                 | 21 (60.5)   | 13           |
|                        | 1 = Farmer      | 195                | 123 (61.1)  | 72           |
|                        | 2 = Merchant    | 82                 | 57 (28.36)  | 25           |

**Accelerated failure time models**

After checking the significance of categorical predictors by K-M survival function and the log-rank test, different AFT models were fitted. Among the Log-normal, Log-logistic, Exponential, Gamma and Weibull distributions, the Weibull AFT model was selected as the AIC and BIC were the smallest as showed in Table 4. Hence, the Weibull AFT model with a combination of covariates was used as the baseline hazard distribution of the parametric shared frailty model.
The overall goodness of fit for the AFT model was checked by the Cox-Snell residual plots [43]. As Fig. 3 showed, the line related to the Cox-Snell residual of the Weibull AFT model was the nearest to 45° straight lines of the origin when compared to that of exponential, log-normal and log-logistic models (Fig. 3a-d).

**Parametric shared frailty models**

Gamma distribution is the most common and widely used in literature for determining the frailty effect [40, 45]. Accordingly, the Gamma frailty and Weibull Gamma shared frailty model was fitted to select the best model for this study using hospitals as random (frailty). The AIC for the Weibull gamma shared frailty (−6003.79) was smaller than the AIC for the Weibull AFT (−600.58) and Gamma shared frailty (−592.53) models. The frailty for the selected model was estimated to be 1.2397 (chi-square = 34.07, df = 1, p-value = 0.0001) which indicated existence of unobserved heterogeneity between the hospitals and that the frailty component in the model was important. Hence the recovery time of MDR-TB patients in Amhara region, Ethiopia was carried out by the Weibull Gamma shared frailty model with hospitals as a clustering effect. Table 5 depicted the result of Weibull Gamma shared frailty model of parameter estimates, time ratio, standard error of estimates, z-value, p-values and 95% CI.

After adjusting other covariates, the hazard of attaining immunological recovery of patients that had extrapulmonary MDR-TB (ϕ=1.0546, 95% CI: 1.0296–1.0802), clinical complication (ϕ=1.0405, 95% CI: 1.025–1.056), comorbidity (ϕ=1.0814, 95% CI: 1.0526–1.1109), male (ϕ=1.037287, 95% CL: 1.0146–1.043) and having poor adherence (ϕ = 1.0316, 95%CI: 1.0063–1.0575) were 5.46, 4.05, 8.14, 3.73, and 3.16% higher than that of the counter groups.
Discussion
In this study, the 311 MDR-TB patients in Amhara region, Ethiopia were assessed, of which, 110 (35.4%) were censored and 201 (64.6%) were cured of MDR-TB. Overall, the median recovery time of MDR-TB patients in Amhara region was 21 months, which means that the recovery time of patients is within the recommended treatment interval of 18 to 24 months or longer [10] given close monitoring of patients while taking these drugs.

The Weibull AFT model had the smallest AIC compared to that of Log-normal, Log-logistic, Exponential, Gamma and gamma AFT models. After selected the Weibull AFT model, Weibull Gamma shared frailty model was well fitted to the data onto Weibull AFT, Gamma frailty and Weibull Gamma shared frailty models. This is because of the model lifetime, Weibull distribution is mostly used in the literature as the Hazard rate for Weibull distribution is a monotone function [31, 46–48] and the Newton–Raphson procedure is used to maximize the best-unbiased estimator for the model parameters [39]. Thus, the MDR-TB patients in Amhara region, Ethiopia were analyzed using the Weibull gamma shared frailty model. This study identified a number of factors associated with the recovery time of MDR-TB patients in Amhara region, Ethiopia taking the hospital’s clustering effect. The clustering effect was significant (p-value < 0.001) in Weibull-gamma shared frailty model that indicates heterogeneity between hospitals and patients within the same hospital share similar risk factors on cure time. This showed that the correlation between hospitals cannot be ignored and the clustering effect was important in modeling the hazard function. Exploring patients increased insight into patient handling, and gave valuable information about providing attention and quality services to patients. The study showed that extrapulmonary MDR-TB patients had longer recovery time than

| Covariates            | Categories  | Median | Log Rank test | Breslow test |
|-----------------------|-------------|--------|---------------|--------------|
|                       |             |        | Chi-square    | Df | P-value | Chi-square | Df | P-value |
| Sex                   | Female      | 20     | 33.18         | 1  | < 0.001 | 27.17      | 1  | < 0.001 |
|                       | Male        | 21     | 0.39          | 3  | 0.94    | 1.58       | 3  | 0.66    |
| Marital status        | Married     | 21     | 0.05          | 1  | 0.94    | 0.05       | 1  | 0.82    |
|                       | Single      | 21     | 6.91          | 3  | 0.07    | 9.07       | 3  | < 0.03  |
|                       | Divorce     | 21     | 27.52         | 1  | < 0.001 | 23.1       | 1  | < 0.001 |
|                       | Widowed     | 19     | 21            |    |         |            |    |         |
| Residence             | Urban       | 21     | 15.44         | 1  | < 0.001 | 11.67      | 1  | < 0.001 |
|                       | Rural       | 21     | 55.14         | 1  | < 0.001 | 44.11      | 1  | < 0.001 |
| Education             | Not Educated| 21     | 23.03         | 1  | < 0.001 | 12.05      | 1  | < 0.001 |
|                       | Primary     | 21     | 37.59         | 2  | < 0.001 | 21.77      | 1  | < 0.001 |
|                       | Second      | 20     | 38.38         | 1  | < 0.001 | 19.77      | 1  | < 0.001 |
|                       | Above       | 20     | 45.65         | 1  | < 0.001 | 29.7       | 1  | < 0.001 |
| Disease category      | Pulmonary   | 21     | 27.52         | 1  | < 0.001 | 23.1       | 1  | < 0.001 |
|                       | Extra pull  | 24     | 21            |    |         |            |    |         |
that of seamier pulmonary MDR-TB patients in Amhara region, Ethiopia, and it was supported by Parmar et al study [49].

Male MDR-TB patients were associated with a high likelihood of experiencing unsuccessful treatment outcomes. Several other studies have similarly shown that male MDR-TB patients tend to have longer recovery time [49–51]. This means that male MDR-TB patients were experiencing longer recovery time than that of female MDR-TB patients. This might be related to higher tendencies towards alcohol and drug abuse, and interruption of their medication as a male has high economical consequences than that of females, while the biochemical, behavioral and socioeconomic determinants of males need further study. The MDR-TB patients with co-morbidity and clinical complication also experienced longer recovery time than that of the control groups. This result is in line with the previous findings in Ethiopia [52, 53] and in India [49]. Furthermore, TB with clinical complication/co-infection often associates with poor response to TB-medications due to nutrition absorption of anti-tuberculosis drugs, overlap toxicity in ART patients and/or due to the risk of immune reconstitution inflammatory syndrome [54]. Unlike this finding HIV status was not a statistically significant predictor variable. This difference may arise from the sample size difference, model difference, as well as the nature of the population. The results of this study revealed that adherence was a significant factor of recovery time of MDR-TB patients in Amhara Region, Ethiopia. These results showed that MDR-TB patients with poor adherence had longer recovery time than MDR-TB patients who had good adherence. This finding was similar with a study [49, 55, 56]. The results of this study suggested that the interaction of smoking status with previous drug use history was a significant predictive factor of time to cure in the Amhara region. This shows that smokers with previously drug user history take more time to cure than patients

| Table 4 Model selection based on IC from AFT models |
|---------------------------------------------------|
| IC       | Exponential | Weibull | Log-normal | Log-logistic | Best model |
| AIC      | 496.83      | -600.58 | -598.52    | -561.97      | Weibull    |
| BIC      | 526.75      | -566.92 | -564.86    | -528.31      | Weibull    |

AIC Akaike’s Information Criteria, BIC Bayesian Information Criteria, IC Information Criteria

Fig. 3 The Cox-Snell residual plots for (a) log-logistic (b) Weibull (c) lognormal (d) exponential distribution
Table 5 Multivariable analysis using Weibull shared Gamma frailty model for the recovery time of MDR-TB patients in Amhara region, Ethiopia

| Variables                  | Categories            | Coeff. T. R(ϕ) | Std Error | Z-value | P > | 95% CI for ϕ |
|----------------------------|-----------------------|----------------|-----------|---------|-----|--------------|
| MDR-TB type                | Extra pulmonary       | 0.05           | 1.055     | 0.013   | 4.34| < 0.001      | 1.03 | 1.08        |
|                           | Pulmonary*            |                |           |         |     |              |      |             |
| Clinical Complication      | Yes                   | 0.04           | 1.041     | 0.008   | 5.23| < 0.001      | 1.025| 1.065       |
|                           | No*                   |                |           |         |     |              |      |             |
| Adherence                  | Fair                  | 0.02           | 1.021     | 0.019   | 1.09| 0.28         | 0.984| 1.059       |
|                           | Poor                  | 0.031          | 1.032     | 0.013   | 2.46| < 0.01       | 1.006| 1.058       |
|                           | Good*                 |                |           |         |     |              |      |             |
| Co morbidities             | Yes                   | 0.078          | 1.081     | 0.015   | 5.69| < 0.001      | 1.053| 1.111       |
|                           | No*                   |                |           |         |     |              |      |             |
| HIV status                 | Positive              | -0.007         | 0.993     | 0.014   | -0.51| 0.613        | 0.966| 1.021       |
|                           | Negative*             |                |           |         |     |              |      |             |
| Sex                        | Male                  | 0.028          | 1.029     | 0.007   | 4.02| < 0.001      | 1.015| 1.043       |
|                           | Female*               |                |           |         |     |              |      |             |
| Previous drug use history  | Yes                   | -0.009         | 0.992     | 0.008   | -1.06| 0.291        | 0.976| 1.007       |
|                           | No*                   |                |           |         |     |              |      |             |
| Smoking status             | Smoker                | -0.051         | 0.95      | 0.02    | -2.43| 0.015        | 0.9125| 0.99        |
|                           | Nonsmoker*            |                |           |         |     |              |      |             |
| Sex by HIV status          | Male HIV positive     | 0.037          | 1.038     | 0.019   | 2.03| 0.042        | 1.001| 1.075       |
|                           | Female HIV negative+  |                |           |         |     |              |      |             |
| Drug use history by smoker | Drug used with smoker | 0.071          | 1.073     | 0.025   | 3.09| 0.002        | 1.026| 1.122       |
|                           | Nondrug use with nonsmoker* |            |           |         |     |              |      |             |
| Cons                      |                       | 2.99           | 19.879    | 0.136   | 438.53| < 0.001      | 19.615| 20.146      |
| Ln(λ)                     |                       | 3.722          | 0.106     | 35.02   | < 0.001| 3.514        | 3.931|
| Ln(θ)                     |                       | -3.986         | 1.775     | -2.25   | 0.03| -7.464       | -0.508|
| λ                         |                       | 41.364         | 4.397     | 33.585  | 50.946|
| γ                         |                       | 0.024          | 0.003     | 0.02    | 0.03|
| θ                         |                       | 1.24           | 0.3608    | 0.695   | 2.212|

γ Shape parameter, λ Scale parameter, θ Frailty parameter, * reference category, CI Confidence interval, Std Error Standard Error, ϕ Time ratio: the ratio for groups of median time to recovery

having no previous history of smoking. This was consistent with the study done by Kuaban et al [57].

Limitations
The limitations of this study were the small sample of participants. There are also a variety of predictors, an assessment of which was beyond the scope of this study, but that might influence the MDR-TB patients. Thus, researchers may consider other predictors to identify important predictors of BDR-TB patients that delay cured time.

Conclusion
This study aimed to investigate the major factors that affect recovery time of MDR-TB patients in Amhara region, Ethiopia. The study revealed that the median recovery time of MDR-TB patients in Amhara region is 21 months, and that close monitoring of patients is required while they take these drugs. In general, it has been concluded that this has the undesirable consequences of economic, social and political weaknesses, shortage of health facilities and access to hospitals in the region. Thus, the regional and federal Government of Ethiopia need to take immediate steps to address the causes of long recovery time of MDR-TB patients in Amhara region, Ethiopia. This may for example involve, strengthening the counseling of patients and family members as well as the use of digital tools in monitoring treatment adherence to decrease the likelihood of recovery time. In addition, there should be robust early case detection and proper treatment of drug-susceptible MDR-TB to shorten the recovery time of MDR-TB in accordance with WHO guidelines.
Abbreviations
ADIS: Acquired Deficiency Immune Syndrome; AFT: Accelerated Failure Time; AIC: Akaike Information Criteria; ART: Antiretroviral therapy; CI: Confidence Interval; CSA: Central statistical agency; HRV: Human Immunodeficiency Virus; MDR-TB: Multidrug resistance tuberculosis; PH: Proportional Hazard; TB: Tuberculosis; WHO: World Health Organization; XDR: Extensively drug-resistant

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Availability of data and materials
The data set supporting the conclusions of this article is held by the authors and the de-recognized data may be made available if a unique request.

Authors’ contributions
YA wrote the proposal, analyzed the data and manuscript writing. DL accredited the proposal with revisions, analysis the data and manuscript writing. Both YA and DL read and approved the very last manuscript.

Ethics approval and consent to participate
The study was ethically cleared by the Research and Ethics Review Board of Bahir Dar University. The letter of support written by the Research and Ethics Review Board was submitted to the selected hospital managers. Prior to data collection, written consent was obtained from the heads of each hospital and confidentiality was maintained by coding from data collection to analysis.

Consent for publication
Not Applicable.

Competing interests
The authors declare that they have no competing interests.

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Reference
1. WHO. Global tuberculosis report 2017. Geneva: World Health Organization; 2017. Licence: CC BY-NC-SA 3.0 IGO
2. WHO. Multidrug and extensively drug-resistant TB (M/XDR-TB) 2010 Global Report On Surveillance And Response 2010.
3. Faustini A, Hall AJ, Perucci CA. Risk factors for multidrug resistant tuberculosis in Europe: a systematic review. Thorax. 2006;61(2):158–63.
4. Kundu D, Sharma N, Chadha S, Lacroix S, Awungatse G, Jiang L, Asaria M. Analysis of multi drug resistant tuberculosis (MDR-TB) financial protection policy: MDR-TB health insurance schemes, in Chhattisgarh state, India. Health Econ Rev. 2018;8(1):3.
5. Singh JA, Upshur R, Padayatchi N. XDR-TB in South Africa: no time for denial or complacency. PLoS Med. 2007;4(1):e50.
6. Koul A, Arnout E, Lounini N, Guillermont J, Andries K. The challenge of new drug discovery for tuberculosis. Nature. 2011;469(7311):483.
7. Assefa D, Seyoum B, Olija L. Determinants of multidrug-resistant tuberculosis in Addis Ababa, Ethiopia. Infect Drug Resist. 2017;10:209.
8. Eldholm V, Balloux F. Antimicrobial resistance in mycobacterium tuberculosis: the odd one out. Trends Microbiol. 2016;24(8):637–48.
9. Getachew T, Bayray A, Weldearegay B. Survival and predictors of mortality among patients under multi-drug resistant tuberculosis treatment in Ethiopia: st. Peter’s specialized tuberculosis hospital, ethiopia. Int J Pharm Sci Res. 2013;4(2):776.
10. Falzon D, Schünenmann HJ, Haraux E, González-Angulo L, Lienhardt C, Jaramillo E, Weyer K. World Health Organization treatment guidelines for drug-resistant tuberculosis, 2016 update. Eur Respir J. 2017;49(3):1602308.
11. Espinal MA, Lazaro A, Simonsen L, Bouhabel F, Kim SJ, Reniero A, Hoffner S, Rieder HL, Binkin N, Dye C. Global trends in resistance to antituberculosis drugs. N Engl J Med. 2001;344(17):1294–303.
12. Tyrell F, Stafford C, Yakrus M, Youngblood M, Hill A, Johnston S. Trends in testing for mycobacterium tuberculosis complex from US public health laboratories, 2009–2013. Public Health Rep. 2017;132(1):56–64.
13. Tacconelli E, Cataldo M, Dancer S, De Angelis G, Falcone M, Frank U, Kählmeter G, Pan A, Petrosillo N, Rodríguez-Baño J. ESCMID guidelines for the management of the infection control measures to reduce transmission of multidrug-resistant gram-negative bacteria in hospitalized patients. Clin Microbiol Infect. 2014;201:55.
14. Alabaih K, Al Aliola S, Al Banyan E, Al Shaalan M, Al Johani S. Characteristics and risk factors of hospital acquired-methicillin-resistant Staphylococcus aureus (HA-MRSA) infection of pediatric patients in a tertiary care hospital in Riyadh, Saudi Arabia. Int J Pediatr Adolesc Med. 2016;3(2):71–7.
15. Tara R, Das P, Kumar D. Recurrent challenges for clinicians: emergence of methicillin-resistant Staphylococcus aureus, vancomycin resistance, and current treatment options. J Lab Phys. 2013;5(2):71.
16. Fair R, Tor Y. Antibiotics and bacterial resistance in the 21st century. Perspect Med Chem. 2014;6:25–64. https://doi.org/10.4137/PMC.S14459.
17. Agerie NW. Determinants of smallholder rural farm households’ participation in small scale irrigation and its effect on income in North Gondar zone: a cross-sectional approach (evidence from Dembia Woreda). Ethiopia: Mekelle University, 2013.
18. Negus D, Lingerew W, Beyene B, Tamiru M, Melaku M. Prevalence of multi drug resistant tuberculosis among presumptive multi drug resistant tuberculosis cases in Amhara National Regional State, Ethiopia. J Mycobac Dis. 2014;4(152):2161–1068.1000152.
19. Cochrane WG. Sampling Techniques. 3rd ed. New York: Wiley, 1977.
20. Fan J. Local polynomial modelling and its applications: monographs on statistics and applied probability 66. New York: Springer; 1992. p. 527–96.
21. van der Meulen A. Life tables and survival analysis. The Hague: Statistics Netherlands; 2012.
22. Babowowska DM. Kaplan-Meier estimate on the plane. Ann Stat. 1988;16(4):1475–89.
23. Brookmeyer R. Median survival time. Wiley StatsRef: Statistics Reference Online 2014.
24. Reid N. Estimating the median survival time. Biometrika. 1983;2(2):273–81.
25. Brueffer C, Vallon-Christersson J, Grabau D, Ehinger A, Häkkinen J, Hegardt C, Malina J, Chen Y, Bendahl P-O, Manjer J. Clinical value of RNA sequencing–based classifiers for prediction of the five conventional breast Cancer biomarkers: a report from the population-based multicenter Sweden Cancerome analysis network—breast initiative. JCO Precis Oncol. 2018;2:1–18.
26. Harrington DP, Fleming TR. A class of rank test procedures for censored survival data. Biometrika. 1982;69(3):553–66.
27. Xie J, Liu C. Adjusted Kaplan–Meier estimator and log-rank test with inverse probability of treatment weighting for survival data. Stat Med. 2005;24(20):3089–110.
28. Kleinbaum DG, Klein M. Kaplan-Meier survival curves and the log-rank test. In: Survival analysis. New York: Springer; 2012. p. 55–96.
29. Collett D. Modelling survival data in medical research. New York: Chapman and HallCRC; 2013.
30. Mo D. DR Regression models and life-tables. In: Breakthroughs in statistics. New York: Springer; 1992. p. 527–41.
31. Gutierrez RG. Parametric frailty and shared frailty survival models. Stata J. 2002;2(1):22–44.
32. Pike M. A method of analysis of a certain class of experiments in carcinogenesis. Biometrics. 1966;22(1):142–61.
33. Bennett S. Analysis of survival data by the proportional odds model. Stat 1993;2(1):22–73–7.
34. Van den Berg GJ. Duration models: specification, identification and multiple durations. In: Handbook of econometrics, vol. 5. Germany: Elsevier; 2001. p. 3381–460.
35. Hougaard P. Modelling heterogeneity in survival data. J Appl Probab. 1991; 28(3):695–701.
36. Aalen O, Borgan O, Geising H. Survival and event history analysis: a process point of view. New York: Springer Science & Business Media; 2008.
37. Hougaard P. Survival models for heterogeneous populations derived from stable distributions. Biometrika. 1986;73(3):374–84.
38. Duchateau L, Janssen P. The frailty Model. New York: Springer-Verlega; 2008.
39. Askin OE, Inan D, Buyuklu AH. Parameter Estimation of Shared Frailty Models based on Particle Swarm Optimization. Int J Stat Probab. 2017;6(1):48–58.
40. Akaike H. A new look at the statistical model identification. IEEE Trans Autom Control. 1974;19(6):716–23.
41. Cox DR, Snell EJ. A general definition of residuals. J R Stat Soc Ser B Methodol. 1968;30(2):248–75.
42. Miller Jr, R.G. Survival Analysis. Vol. 66. Chichaster: Wily; 2011.