Frailty in hemodialysis and prediction of poor short-term outcome: mortality, hospitalization and visits to hospital emergency services

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ABSTRACT

Background: Frailty is an aging-associated state of increased vulnerability, which raises the risk of adverse outcomes. Chronic kidney disease is associated with higher prevalence of frailty. Our aim was to estimate frailty prevalence in a hemodialysis population and its influence on short-term outcomes.

Design: Observational prospective longitudinal study of 277 prevalent hemodialysis patients. Frailty was estimated through the Edmonton Frail Scale (EFS). Demographic and clinical data, comorbidity index, and laboratory parameters were recorded. A 29-month follow-up was conducted on mortality, including hospitalization, and visits to hospital emergency services in the first 12 months of this period.

Results: According to the EFS, 82 patients (29.6%) were frail, 53 (19.1%) were vulnerable, and 142 (51.3%) were non-frail. During follow-up, 58.5% frail patients, 30.2% vulnerable, and 16.2% non-frail ones died (p < .005). In the analysis of survival using an adjusted Cox model, a higher hazard of mortality was observed in frail than in non-frail patients (HR 2.34; 95% CI 1.39–3.95; p = .001). During follow-up the hospitalization rate was 852 episodes/1000 patient-years for frail patients, 784 episodes/1000 patient-years for vulnerable patients, and 417 episodes/1000 patient-years for non-frail patients (p = .0005). The incidence ratio of visits to emergency services was 3216, 1735, and 1545 visits/1000 patient-years for each group (p < .001).

Conclusions: Hemodialysis patients present high frailty prevalence. Frailty is associated with poor short-term outcomes and higher rates of mortality, visits to hospital emergency services, and hospitalization.
Index, or through subjective scales like the Canadian Frailty Scale [11–13]. However, there is less experience in the use of multi-domain scales, like the Edmonton Frail Scale (EFS), in hemodialysis.

Our objective was to estimate frailty prevalence in a hemodialysis population using the EFS [14] – a simple assessment tool comprising eleven items focusing on different frailty dimensions, which can be applied in the clinical practice – and to evaluate the association with demographic, clinical, and laboratory variables, as well as, with poor short-term outcome assessed through visits to hospital emergency services, hospitalization episodes, and death.

Materials and methods

Design

We conducted a prospective, observational, and longitudinal study with patients on the hemodialysis program in the South Healthcare Area of Gran Canaria, Spain, which corresponds to a population of 375,000. To estimate frailty prevalence, a cross-section was established for prevalent patients in October 2016. Inclusion criteria were: patient older than 18 years, more than 3-month prevalence in hemodialysis and ability to understand the information provided and to sign an informed consent form. Exclusion criteria were: patient with an active neoplastic or infectious disease or hospitalized in the previous 3-month period due to infectious, cardiovascular, or surgical complications. Demographic and clinical data were collected through the electronic medical records. Baseline complete laboratory tests were conducted on mid-week pre-dialysis blood extraction samples using the standard laboratory determination methods in our center. The modified Charlson comorbidity index [15,16] was calculated for all patients included in the study.

Patients

Out of 294 available patients, 17 were excluded because they met the exclusion criteria, resulting in a final number of 277 patients included.

Frailty

To estimate frailty prevalence, patients were administered the EFS, which comprises 11 items distributed into nine domains: cognitive (evaluated through the clock-drawing test), general health status, dependence, social support, medication, nutrition, depression, sphincter continence, and a physical test consisting of standing up and walking. Every item can be scored between 0 and 2, so that the global score varies from 0 to 17. Scores 0–5 correspond to non-frail, 6–7 to vulnerable, 8–9 to mild frail, 10–11 to moderate frail, and 12–17 to severe frail. For the statistical analysis, all frailty categories were grouped into frail: scores 8–17.

Follow up

A one-year follow-up was conducted, where all visits to hospital emergency services were recorded through the electronic medical records. Additionally, all hospital admissions, their cause and duration, were recorded. A follow-up of patients was conducted until April 2019 (29 months), recording all deaths and any other reason for abandoning the study before the end of the follow-up period.

Ethics

The study was conducted in accordance with the Declaration of Helsinki and was approved by the Clinical Research Ethics Committee in our Center. All patients signed written informed consent before participating in the study.

Statistical analysis

Because the data did not follow a normal distribution, they were expressed as median and interquartile ranges (percentile 25 and 75, IQR) for continuous variables or as absolute frequency and percentages for qualitative variables. For comparison of qualitative variables between two or more groups, the Chi-squared test or the Fisher’s exact test were used, depending on data distribution. For comparison of continuous variables between two or more groups, the Mann–Whitney U-test or the Kruskal–Wallis test were used, as applicable. For all qualitative and continuous variables we also calculated the \( p \) values for the trend between groups using the appropriate test. All statistical tests were considered bilateral and significance was considered for \( p \) values lower than .05.

Survival times of patients under different frailty conditions were evaluated by a longitudinal cohort study. The Kaplan–Meier method was used to examine crude survival in the three groups defined by their frailty status (frail, vulnerable, and non-frail). Cox proportional hazards regression was applied firstly to estimate unadjusted hazard ratios (HRs) in the three groups. Next, multidimensional Cox regression was
used to adjust for possible confounders. This was performed by entering all the variables potentially associated with survival into the model: Edmonton Frailty, age, sex, diabetes mellitus, months on dialysis, Charlson comorbidity index without age, body mass index (BMI), hemoglobin, albumin, prealbumin, T cholesterol, triglycerides, C reactive protein, uric acid, calcium, phosphate, intact parathyroid hormone (iPTH), creatine kinase, potassium, and creatinine. After using the backward selection method, only the significant variables remained: Edmonton frailty, Charlson comorbidity index without age, BMI, albumin, and creatine kinase. The Cox modeling results were summarized with HRs for each variable, 95% confidence intervals, and associated p values.

A negative binomial regression model was used to model both the number of hospital admissions and the number of emergency visits during the period of patient follow-up as a linear function of frailty status and other possible influential variables. This model was used as an alternative to Poisson regression model for count data due to its ability to adapt to situations in which the average number of events and their variance are different, as in this case. Because events were counted over different time intervals for different individuals, duration of the follow-up period was included in the model as an offset variable. Results of the negative binomial regression models were summarized as incidence rate ratios for each variable, 95% confidence interval and corresponding p values. As in the case of survival times, the negative binomial model was used first to estimate unadjusted incidence rate ratios for each frailty level (considering non-fragile as the reference level), and then a stepwise selection method was used to include possible confounders in the model and calculate adjusted incidence rate ratios.

In all cases, hypothesis testing was considered significant when the corresponding p values were less than .05.

All statistical analyses were conducted with statistical software R version 3.5.3 (R Core Team (2019). R: A language and environment for statistical computing. R Foundation for Statistical Computing, Vienna, Austria. URL https://www.R-project.org/).

**Results**

The study included 277 patients of 65 years median age (IQR 53–73): 182 men (65.7%) and 95 women (34.3%), with 34.6 months median time in dialysis (IQR 15.6–74.5); 159 patients were diabetic (57.4%). The most frequent etiology of renal disease was diabetic nephropathy (41.5%) followed by nephropathy of vascular origin (13%), interstitial nephropathy (8.3%), glomerular nephropathy (7.9%), ADPKD (7.9%), and nephropathy of other or unknown origin (the remaining 21.4%). The median Charlson index was 6 (IQR 5–8).

According to the EFS, 82 patients (29.6%) were frail, 53 patients (19.1%) were vulnerable, and 142 patients (51.3%) were non-frail. Among frail patients, 48.8% were mild frail, 28% were moderate frail, and 23.2% were severe frail. Figure 1 shows these results.

Table 1 shows demographic and laboratory variables for the different frailty groups. An association was found between frailty and certain variables often associated with poor prognosis, such as age, diabetes, high Charlson comorbidity index or female sex. Regarding laboratory parameters, we would like to highlight the association between frailty and lower hemoglobin, albumin, prealbumin, uric acid, phosphate, creatinine, and creatine kinase counts.

In the patient follow-up to April 2019, the mean follow-up time was 22 ± 9 months. In total, 145 patients (52.3%) completed the follow-up period, 87 (31.4%) died, 42 (15.2%) received a kidney transplant, and 3 (1.1%) were lost to follow-up after transferring to another medical center. The number of deaths was 48 patients in the frail group (58.5%), 16 in the vulnerable group (30.2%) and 23 in the non-frail group (16.2%). Figure 2 shows the Kaplan–Meier survival curves, with lower survival of frail than non-frail patients (p < .001). In the Cox model of patient survival, the unadjusted hazard rate of mortality for frailty, using the non-fragile group as a reference, was 1.8; 95% CI 0.94–3.5; p = .075 and 4.1; 95% CI 2.51–6.8; p < .001 for vulnerable and frail patients, respectively. When adjusted by the other significant variables for mortality, HR of frail compared to non-frail patients was 2.34; 95% CI 1.39–3.95; p = .001. Vulnerable patients did not show a significantly higher hazard than non-frail patients. The other

![Figure 1](https://example.com/figure1.png)
significant variables maintained in the model were the Charlson index without age, BMI, albumin level, and CPK (Figure 3).

Excluding hospitalization for kidney transplant, the rates of hospitalization/1000 patient-years were 852, 784, and 417 episodes for frail, vulnerable and non-frail patients, respectively. To compare the incidence ratio of hospitalization between the three groups, a negative binomial regression model was used. Table 2 shows that the rate of hospitalization is significantly higher in the frail and vulnerable patients than in the non-frail patients, both in the unadjusted model and the model adjusted for other significant variables. Table 3 shows the data corresponding to hospitalizations in the three groups. It can be observed that the number of hospitalization days per patient and year, and the mean duration of hospital stay per hospitalization episode, were higher in frail patients. Regarding hospitalization causes, non-frail patients showed higher tendency to hospitalization due to complications with dialysis access and surgical interventions; while frail patients showed higher tendency to hospitalization due to infectious or cardiovascular complications. Additionally, bone fractures as the cause of hospitalization appeared in this group, but not in the others.

The rates of visits to emergency services were 3216, 1735, and 1545 visits/1000 patient-years for frail, vulnerable and non-frail patients, respectively. To compare the rate of visits to hospital emergency services between the three groups, a univariate negative

![Figure 2. Kaplan–Meier survival curves by frailty status. Log-Rank test p values <.001.](image1)

![Figure 3. Multivariable Cox proportional hazards models of the association of frailty and mortality adjusted by Charlson Comorbidity Index without age, body mass index, serum albumin, and creatin kinase.](image2)

| Table 1. Edmonton frail scale (EFS). |
|--------------------------------------|
| **Non-frail** | **Vulnerable** | **Frail** | **p Value for trend** |
| Age | 62 (50–69) | 65 (54–76) | 71 (63–78) | <.001 |
| % Sex (male) | 72.5 | 67.9 | 52.4 | .003 |
| % Diabetes | 43 | 62.3 | 79.3 | .001 |
| Charlson comorbidity index | 6 (4–7) | 7 (4.5–9) | 8 (6–10) | <.001 |
| Months on dialysis | 29 (15–67) | 33 (17–78) | 50 (21–82) | .047 |
| Body mass index | 27 (24–30) | 27 (24–31) | 26 (24–30) | .776 |
| Hemoglobin, g/dL | 11.7 (11–12.3) | 11.3 (10.6–11.9) | 11.1 (10.2–12) | .001 |
| Glucose, mg/dL | 109 (95–146) | 114 (95–168) | 138 (108–194) | .001 |
| Albumin, g/dL | 3.7 (3.4–3.8) | 3.5 (3.3–3.8) | 3.5 (3.2–3.8) | .001 |
| Prealbumin, mg/dL | 28 (24–33) | 26 (21–31) | 26 (20–30) | <.001 |
| Uric acid, mg/dL | 6.4 (5.4–7.3) | 6.5 (5.3–7.1) | 5.6 (5–6.3) | .001 |
| Calcium, mg/dL | 8.8 (8.2–9.3) | 8.9 (8.5–9.2) | 8.8 (8.4–9.3) | .550 |
| Phosphate, mg/dL | 4.4 (3.7–5.4) | 4.5 (3.5–5.4) | 4 (3.4–4.8) | .027 |
| iPTH, pg/mL | 277 (182–414) | 225 (174–458) | 272 (173–391) | .672 |
| Creatin kinase, U/L | 91 (58–143) | 67 (50–107) | 50 (29–82) | <.001 |
| Creatinine, mg/dL | 8.5 (6.4–10.4) | 7.5 (6.1–9.7) | 7.2 (5.6–8.3) | .001 |
| T cholesterol, mg/dL | 147 (121–175) | 141 (124–156) | 135 (112–163) | .033 |
| Triglycerides, mg/dL | 130 (92–191) | 126 (95–181) | 113 (80–181) | .121 |
| Potassium, mM/L | 5.3 (4.6–6.1) | 5.5 (4.6–6.4) | 5.3 (4.7–5.7) | .497 |
| C reactive protein, mg/dL | 0.41 (0.19–0.99) | 0.61 (0.24–1.12) | 0.48 (0.22–1.09) | .377 |
binomial regression model and a multivariate model adjusted for other significant explanatory variables were used.

Table 4 shows the results, in which the rate of visits to hospital emergency services among frail patients is almost twice that of non-frail patients and the difference is statistically significant after adjusting for other variables ($p = .0002$).

**Discussion**

While there is no consensus on the definition of frailty, it is currently recognized that frailty is an age-associated status of reduced resilience and increased vulnerability to stressful situations, characterized by weakness and lower biological functional reserves, which entails higher risk of unfavorable outcomes toward disability, hospitalization, and death [1,2,17]. Although it is closely related to sarcopenia, dependence, comorbidity, and disability, and these situations influence each other, they do not always coexist [18].

Although frailty was initially defined as associated with aging, situations have been described, such as chronic kidney disease, which may not only increase its prevalence at advanced ages, but also result in earlier onset [5–7,19].

**Table 2.** Negative binomial regression hospitalization.

|            | IRR    | Pr(>lzl) |
|------------|--------|----------|
| **Unadjusted** |        |          |
| Intercept   | 0.03 [0.03; 0.05] | $p < .0001$ |
| Vulnerable  | 1.94 [1.20; 3.13] | .0056    |
| Frail       | 2.09 [1.38; 3.18] | .0005    |
| **Adjusted** |        |          |
| Intercept   | 0.05 [0.01; 0.29] | .0007    |
| Vulnerable  | 1.82 [1.13; 2.92] | .0124    |
| Frail       | 1.78 [1.15; 2.77] | .0094    |
| Charlson CI without age | 1.16 [1.06; 1.27] | .0014 |
| Albumin     | 0.61 [0.39; 0.96] | .0283    |
| Phosphate   | 1.18 [1.03; 1.34] | .0100    |

Large differences in the prevalence of frailty in hemodialysis patients found in previous studies may be due to several reasons, such as differences in the study populations in terms of age, morbidity, and inclusion of prevalent or incident patients. However, they may also be due to differences in the methods used to detect frailty. In general, three frailty-detection tests are used: those based on the description of a frail phenotype through physical tests, such as the one described by Fried [4]; those based on subjective assessment scales, such as the Frailty Score developed by Rockwood in Canada [20], and multi-domain tools that explore various frailty dimensions such as cognitive state, degree of dependence, psychological dimension, social support, or physical dimension, e.g., the Groningen Frailty Indicator [21], the Tilburg Frailty Indicator [22], and the EFS [14].

Most studies conducted on hemodialysis patients have used the Fried Frail Phenotype, observing frailty prevalence of 21.9–73% [11,12,23,24].

Studies on hemodialysis in which subjective scales were used to assess frailty described a prevalence of 19.6–26% [13,25].

Multi-domain tests assessing a range of frailty dimensions have been less studied in hemodialysis. Van Loon observed 67% frailty prevalence using the...
Groningen Frailty Indicator, although the study was conducted on an advanced-age population of 123 incident hemodialysis patients, all older than 65 years [26]. We have recently published our data for prevalent patients on hemodialysis measuring frailty using the Fried Phenotype Frailty Index and the EFS, where the proportion of frail patients was 41.2% when measured with the Fried criteria and 29.6% when measured with the EFS [27].

We chose the EFS because it can be easily administered in daily clinical practice and has been validated for different populations [9,28–30]. Using this test, we found 29.6% frailty prevalence in our population, which was close to the data reported with other methods in previous studies. Few studies have been published on the use of this test in hemodialysis patients. We found only two small studies: a study by Chao in Taiwan, including only 46 patients and reporting 43.6% frailty prevalence with the EFS, and a study by Orlandi, in Brazil, including 60 hemodialysis patients, all older than 60 years, reporting 38.3% frailty prevalence [25,31]. Our study is therefore the most extensive to date in which the EFS is applied to hemodialysis patients.

As shown in Table 1, frailty in hemodialysis patients is associated with demographic and clinical factors, which are usually associated with poor prognosis, such as older age, diabetes mellitus, and higher Charlson comorbidity index [32,33]. An association was also found between frailty and certain laboratory parameters such as lower hemoglobin, albumin, and prealbumin [34,35]. However, the association with laboratory parameters of bone and mineral metabolism and lipids is less clear. We failed to find an association with the inflammation marker C-reactive protein [36,37]. The association between frailty and serum creatinine and creatine kinase may be accounted for by lower muscle mass due to frailty-associated sarcopenia [38,39].

This relation found between surrogate markers of sarcopenia and frailty is of great interest. Sarcopenia is highly prevalent in hemodialysis and is associated not only with frailty, but also with the protein-energy wasting (PEW) syndrome. This syndrome, also known as malnutrition inflammation complex syndrome (MICS), is associated with poor prognosis [40,41]. We also found an association between frailty and markers of malnutrition (serum albumin and prealbumin). Further studies are needed to determine the relation between frailty and PEW in dialysis and its influence on prognosis.

Our results show a significant relation between frailty and mortality after a 29-month follow-up. In the Cox proportional hazards model, with the addition of other variables significantly related to mortality (age, diabetes mellitus, Charlson comorbidity index without age, and BMI) and analytical parameters associated with bone and mineral metabolism, lipid profile and nutrition parameters (albumin, prealbumin, creatinine, and CPK), the HR of mortality in frail patients compared to non-frail patients was 2.34; 95%CI 1.39–3.95; p = .001. These results agree with others previously published. In a study by McAdams-DeMarco on 146 hemodialysis prevalent patients, frailty was found to be a potent predictor of mortality at three years, when the risk was 2.6 times higher (CI 95%: 1.04–6.49) [12]. Johansen, in a study with 762 hemodialysis patients, observed a 2.1-fold higher risk of death at 2 years (CI 95%: 1.41–3.29) [11]. In both studies, frailty was assessed through the 5-item frailty phenotype described by Fried, including strength measured by a dynamometer and the walking speed test.

Two further studies with larger patient populations have been published: one by Bao on incident patients [24] and the other by Lee on prevalent patients [42]. In both studies, the Fried frail phenotype evaluated through questionnaires was used. Both studies showed independent association between frailty and mortality, with 1.57 (CI 95%: 1.25–1.97) and 2.37 (CI 95%: 1.11–5.02) HR in the first and second study, respectively. Alfaadhel evaluated frailty through the subjective clinical frailty scale in 390 incident hemodialysis patients. In this study, every incremental point on the frailty scale was associated with an increase in the mortality risk with a 1.22 HR (CI 95%: 1.04–1.43) [13]. Our results are in agreement with all of these studies. However, ours is the first study conducted on hemodialysis patients to show an association between mortality and frailty assessed through a multi-domain test that explores multiple aspects of frailty.

The number of hospitalization episodes has also been considered a sign of poor clinical prognosis, which may be associated with frailty. In our study, the percentage of patients admitted to hospital was significantly higher in frail and vulnerable patients than in non-frail patients. We also observed higher rates of hospitalization per 1000 patient-years among frail and vulnerable patients than among non-frail patients. In the cited studies by McAdams-DeMarco [12], Bao et al. [24], and Lee et al. [42], associations were demonstrated between frailty, number of hospitalization episodes, and time until first hospitalization. However, in those studies, frailty was assessed through the Fried Frailty Phenotype; while our study is the first to associate hospitalization with frailty assessed through a multi-domain test like the Edmonton test. Furthermore, in
our study, a higher number of hospitalization episodes and longer mean hospital stay were found for frail patients. However, the mean stay duration was not adjusted for complexity and it might reflect a different patient profile, given that, as previously shown, frail patients are usually older, show higher mortality and their hospitalization causes tend to be different, e.g., cardiovascular and infectious causes are more frequent among frail patients, while vascular access complications and major surgery are more frequent among non-frail patients.

The number of visits to the emergency service is important information for the management of healthcare resources. It has been reported that failure to provide early primary care close to the patient’s home, in frail and multimorbid patients, impairs prevention of complications, and leading to frequent visits to hospital emergency services. This is important, because it may result in saturation of these services, thus preventing resource allocation to acute, potentially curable conditions, and increasing healthcare costs. Furthermore, the time frail patients spend in saturated emergency departments often results in healthcare-related adverse events, mainly of an infectious nature [43-45]. To the best of our knowledge, our study is the first to relate frailty in hemodialysis to the number of visits to hospital emergency services. We found that the rate of annual visits to emergency services in frail patients was twice that of non-frail patients. This difference was significant after adjusting for other variables in a negative binomial regression model. These findings are important because some studies have demonstrated that adequate frailty management through nutritional intervention strategies, physical exercise, and healthcare support, administered in the primary healthcare setting or in the dialysis units, may enhance patients’ baseline situation and eventually prevent poor short-term prognosis [46,47].

The EFS is very simple to administer and can be carried out by hemodialysis nursing staff after minimal training, without the need for specialist medical staff, in less than 15 min. Its association with poor short-term outcome highlights its utility in detecting frail patients, but because it is a multi-domain test, it also gives health professionals information about which areas of frailty are affected. This would allow specialists to administer specific, more complex scales to detect deficits in independence for basic and instrumental activities of daily life, cognition, emotional status, and social support. Detecting at-risk patients for a broader, more specific global geriatric assessment would allow health professionals to identify areas that require higher priority preventive or corrective action by the most appropriate professionals: geriatricians, neurologists, psychologists, physiotherapists, occupational therapists, and primary care and social work teams. The purpose of applying these preventive or corrective measures is to improve patient outcome and avoid complications associated with frailty.

The main strength of our study is that it is the first to analyze the association between frailty in hemodialysis patients, measured through a multi-domain test like the EFS, and poor clinical prognosis of patients. Additionally, we consider that including visits to hospital emergency services as a poor-prognosis variable gives greater interest to the study.

This study has also several limitations. First, it is a single-center study with a limited number of patients. It could also be argued that the EFS has not been validated in the general Spanish population, although many of the items used are commonly applied in Spain in other global geriatric assessment scales to measure dependence, depression, cognitive decline, and comorbidity. The EFS has also been used in a wide range of populations, including others in the Mediterranean that are similar to the Spanish population, e.g., Portuguese and Italian [28,48] and a Spanish version of the test validated in a Spanish speaking population in Colombia has been used [49]. Conducting the study on the prevalent but not the incident patient population may also be considered a limitation, because the frailty situation of patients at the beginning of renal replacement therapy was not known and may have been influenced by the time on therapy. Additionally, frailty was assessed at a particular time point, with no longitudinal follow-up, and therefore possible variations in the state of frailty and their impact on patient prognosis were not detected. However, a large impact on results is not expected, because this was a short-term follow up analysis. It would be interesting to compare the impact of frailty on patient outcome, assessed through a multi-domain test versus the frail phenotype, to determine whether including further frailty dimensions provides new data that could be useful for prognosis. For these reasons, the results of our study provide interesting data, although they should be taken with caution and supported by wider studies comparing the results of various frailty tests.

Our study demonstrated frailty prevalence in hemodialysis, assessed through the EFS, which was consistent with most of the studies where other test types were used. The results also showed that the occurrence of frailty, assessed through a multi-
domain test was associated with poor outcome in hemodialysis patients, evidenced in higher mortality, hospitalization rates, and frequency of visits to hospital emergency services.

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Disclosure statement

No conflict of interest declared by authors.

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