Seizures after Ischemic Stroke: A Matched Multicenter Study

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Objective: The purpose of this study was to identify risk factors for acute symptomatic seizures and post-stroke epilepsy after acute ischemic stroke and evaluate the effects of reperfusion treatment.

Methods: We assessed the risk factors for post-stroke seizures using logistic or Cox regression in a multicenter study, including adults from 8 European referral centers with neuroimaging-confirmed ischemic stroke. We compared the risk of post-stroke seizures between participants with or without reperfusion treatment following propensity score matching to reduce confounding due to treatment selection.

Results: In the overall cohort of 4,229 participants (mean age 71 years, 57% men), a higher risk of acute symptomatic seizures was observed in those with more severe strokes, infarcts located in the posterior cerebral artery territory, and

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post-stroke seizures are related to stroke severity, etiology, and location, whereas an early electroencephalogram was not predictive of epilepsy. We did not find an association of reperfusion treatment with risks of acute symptomatic seizures or post-stroke epilepsy.

**Methods**

**Study Population**

We used data from the SeLECT study of post-stroke seizures, currently involving 8 international cohorts. In brief, we included consecutive individuals aged 18 years or older with neuroimaging-confirmed acute ischemic stroke. We excluded those with a medical history of seizures or epilepsy before stroke, primary hemorrhagic stroke, no evidence of stroke on neuroimaging, transient ischemic attack, potentially epileptogenic comorbidities (ie, intracranial tumor, cerebral venous thrombosis, history of severe traumatic brain injury, and history of brain surgery), or re-infarction during follow-up. The study flow-chart is shown in Fig 1. Follow-up was performed using face-to-face interviews (in Austria, Italy, Portugal, and Switzerland [2]), telephone interviews (in Germany [2], Spain, and Switzerland [2]), medical chart review (in Germany [1]), or telephone screening followed by face-to-face interview (in Switzerland [1]).

Regulatory approval was granted from all local ethical committees. All subjects in the Italian, Spanish, and Portuguese cohort and those having a face-to-face interview in the Swiss cohort gave written informed consent. All subjects evaluated by telephone in the Swiss and German (2) cohorts gave verbal informed consent. According to Swiss and German law, the regional ethical committees exempted these cohorts from requiring written informed consent. The Austrian case-control study was classified as a retrospective service evaluation by the regional ethical committee and informed consent was not required.
Definitions

We used the World Health Organization (WHO) definitions for stroke and ILAE definitions for seizures, which were classified as acute symptomatic (≤7 days post-stroke) or remote symptomatic (spontaneous unprovoked seizures >7 days post-stroke, congruent with post-stroke epilepsy). Acute provoked seizures (eg, due to electrolyte disturbance, induced by medication, etc.) were included as acute symptomatic, however, remote provoked seizures were not considered as post-stroke epilepsy. Stroke severity was estimated using the National Institutes of Health Stroke Scale (NIHSS). Stroke etiology was categorized according to the Trial of Org 10172 in Acute Stroke Treatment (TOAST) classification. Involvement of the cerebral cortex and/or vascular territory by irreversible infarction was determined by best available neuroimaging (computed tomography [CT] or magnetic resonance imaging [MRI]).

We also evaluated a subset of patients who received a routine EEG using the 10–20 system within the first 7 days after stroke onset. We assessed diffuse background activity slowing, focal slowing, interictal epileptiform discharges, lateralized periodic discharges, electrographic seizures, and status epilepticus in the EEG recordings based
on established definitions. The EEG was classified as abnormal if any of the above factors were present.

Missing values (0.8%) in the 4 original SeLECT cohorts were imputed using multiple imputation methods, as previously described (see online supplement for details on missing values). Data in the 4 newly added cohorts was complete and imputation was not necessary.

Our primary outcome was time to post-stroke epilepsy. Our secondary outcome was the occurrence of acute symptomatic seizures after starting reperfusion treatment in participants with acute ischemic stroke.

**Statistical Analysis**

To evaluate variables that are associated with the risk of acute symptomatic seizures, we first used univariable logistic regression and subsequently analyzed all significant (p < 0.05) variables in a multivariable logistic regression model (Table 1). We applied a similar 2-step approach using Cox regression to assess factors associated with time to post-stroke epilepsy (Table 2). We also evaluated the predictive role of early (within the first 7 days following stroke) EEG for post-stroke epilepsy using Cox regression and the two-step approach (Table 3). For this, we selected individuals with available early EEG data after stroke (n = 673).

Comparisons of those who did and did not receive reperfusion treatment are potentially confounded by treatment selection bias. Severe strokes with large vessel occlusion affecting the cerebral hemispheres are more likely to receive reperfusion treatment, and more likely to be associated with subsequent seizures. We used propensity score matching (PSM) to reduce treatment-selection bias and to mimic the characteristics of a randomized controlled trial in regard to comparing treatment groups without the need for additional adjustment using regression methods. PSM is considered a reliable method to achieve an optimal between-group balance in covariates. It provides treatment effects estimates that are closer to the true marginal treatment effect, making it a superior alternative to classical regression particularly when the number of covariates is 5 or more.

First, we compared the baseline characteristics between participants who received reperfusion treatments

| TABLE 1. Logistic Regression Analysis of Predictors Associated with Acute Symptomatic Seizures Following Acute Ischemic Stroke |

|                              | Univariable |              | Multivariable |              |
|------------------------------|-------------|--------------|--------------|--------------|
|                              | OR (95% CI) | p value      | aOR (95% CI) | p value      |
| Age                          | 1.00 (0.98–1) | 0.532       |              |              |
| Sex (male)                   | 1.02 (0.91–1.64) | 0.178       |              |              |
| NIHSS at admission           |              |              |              |              |
| 0–3                          | 0.77 (0.56–1.02) | 0.075       |              |              |
| 4–10                         | 0.86 (0.62–1.15) | 0.315       |              |              |
| ≥11                          | 2.14 (1.58–2.87) | <0.001     | 1.68 (1.22–2.31) | 0.001     |
| Stroke etiology              |              |              |              |              |
| Large-artery atherosclerosis | 1.82 (1.32–2.51) | <0.001     | 1.43 (1.03–1.98) | 0.025     |
| Small-vessel occlusion       | 0.14 (0.05–0.26) | <0.001     | 0.19 (0.08–0.39) | <0.001    |
| Cardioembolic                | 1.09 (0.8–1.48) | 0.553       |              |              |
| Stroke location              |              |              |              |              |
| Middle cerebral artery       | 1.22 (0.89–1.71) | 0.204       |              |              |
| Anterior cerebral artery     | 1.34 (0.71–2.3) | 0.321       |              |              |
| Posterior cerebral artery    | 1.51 (1.1–2.04) | 0.008       | 1.46 (1.07–2.00) | 0.015    |
| Cortical involvement         | 0.02 (1.32–2.39) | <0.001     | 1.32 (0.98–1.79) | 0.069    |
| Reperfusion treatment        | 1.48 (1.09–1.99) | 0.008       | 1.04 (0.75–1.42) | 0.817    |

OR = odds ratio; aOR = adjusted odds ratio; CI = confidence interval; NIHSS = National Institutes of Health Stroke Scale.
and those who did not using chi-square tests for normally distributed data and Wilcoxon rank sum tests for those with a non-normal distribution (Supplementary Table S1). The reported baseline characteristics also included those that were previously identified as independent predictors of seizure risk (severity of stroke, large-artery atherosclerotic etiology, cortical involvement, and middle cerebral artery territory involvement). We also used multivariable Cox proportional hazards analysis to explore covariates associated with post-stroke epilepsy and multivariable logistic regression to identify those associated with acute symptomatic seizures before matching. All predictors were assessed for proportional hazards by statistical and visual inspection of log(−log) plots. Multicollinearity was assessed with the variance inflation factor (VIF). Martingale residuals plots were analyzed to assess for nonlinearity. Assumptions related to the study design and data collection (left censoring, noninformative censoring, and secular trends) were fulfilled.

Second, we selected a subgroup matched on the logit of the propensity score using a greedy nearest neighbor algorithm within a caliper width of 0.2 of the standard deviation of the logit of the propensity score. This method results in less biased estimates compared to other matching algorithms. Matching was performed on all covariates (see Supplementary Table S1). Last, to assess for balance of covariates between groups before and after matching, we visually evaluated density function plots and estimated the standardized mean difference (SMD), an SMD greater than 0.15 being indicative of poor balance.

As a secondary analysis, we repeated the same matching procedure to compare participants who received IV thrombolysis using recombinant tissue plasminogen activator (rtPA) versus those who received no reperfusion treatment (Supplementary Table S2). The rationale was to specifically explore the effects of rtPA on post-stroke seizures rather than

| TABLE 2. Cox Regression Analysis of Factors Associated with Time to Post-Stroke Epilepsy Following Acute Ischemic Stroke |
|---|---|---|---|
| **Univariable** | **Multivariable** |
| **HR (95% CI)** | **p value** | **aHR (95% CI)** | **p value** |
| **Age** | 0.99 (0.99–1.01) | 0.096 | |
| **Sex (male)** | 0.86 (0.66–1.11) | 0.266 | |
| **NIHSS at admission** | | | |
| 0–3 | 0.41 (0.31–0.55) | <0.001 | 0.75 (0.52–1.09) | 0.136 |
| 4–10 | 0.79 (0.59–1.05) | 0.108 | |
| ≥11 | 3.62 (2.79–4.69) | <0.001 | 2.33 (1.69–3.21) | <0.001 |
| **Stroke etiology** | | | |
| Large-artery atherosclerosis | 2.07 (1.52–2.64) | <0.001 | 1.51 (1.13–2.01) | 0.004 |
| Small-vessel occlusion | 0.37 (0.24–0.57) | <0.001 | 0.92 (0.57–1.46) | 0.728 |
| Cardioembolic | 0.98 (0.73–1.30) | 0.901 | |
| **Stroke location** | | | |
| Middle cerebral artery | 2.26 (1.62–3.15) | <0.001 | 1.35 (0.95–1.92) | 0.087 |
| Anterior cerebral artery | 0.98 (0.50–1.91) | 0.957 | |
| Posterior cerebral artery | 1.15 (0.86–1.54) | 0.330 | |
| Cortical involvement | 2.75 (2.07–3.66) | <0.001 | 2.07 (1.54–2.80) | <0.001 |
| ASS | 5.87 (4.11–8.38) | <0.001 | 4.18 (2.90–6.02) | <0.001 |
| Reperfusion treatment | 1.85 (1.41–2.40) | <0.001 | 1.05 (0.78–1.40) | 0.727 |

HR = hazard ratio; aHR = adjusted hazard ratio; ASS = acute symptomatic seizure; CI = confidence interval; NIHSS = National Institutes of Health Stroke Scale.
the effects of reperfusion, per se. As another secondary analysis, we performed the same procedures for individuals who received mechanical thrombectomy only versus those who received no reperfusion treatment. Data in the German (2) cohort did not distinguish between types of reperfusion treatment, so participants receiving reperfusion treatment in this cohort were not included for these secondary analyses.

We compared time to post-stroke epilepsy (primary outcome) between participants with or without reperfusion treatment using the log-rank test and the Kaplan Meier method. The occurrence of acute symptomatic seizures (secondary outcome) was compared using logistic regression, because time to event was <1 week and thus pragmatically classified as binomial.

| Predictors | A: Univariable | B: Multivariable |
|------------|----------------|-----------------|
|            | HR (95% CI)    | p value         | aHR (95% CI)    | p value         |
| Age at stroke | 0.99 (0.97–1.00) | 0.566           |                  |                 |
| Sex (male) | 1.14 (0.67–1.93) | 0.625           |                  |                 |
| NIHSS at admission |  |  |
| 0–3 | 0.36 (0.2–0.66) | 0.001           | 0.55 (0.25–1.21) | 0.140           |
| 4–10 | 0.93 (0.53–1.63) | 0.807           |                  |                 |
| ≥11 | 3.37 (1.99–5.70) | <0.001          | 1.27 (0.65–242) | 0.460           |
| Stroke etiology |  |  |
| Large-artery atherosclerosis | 2.78 (1.62–4.78) | 0.002           | 1.62 (0.90–2.90) | 0.104           |
| Small-vessel occlusion | 0.55 (0.28–1.10) | 0.093           |                  |                 |
| Cardioembolic | 0.63 (0.33–1.23) | 0.182           |                  |                 |
| Stroke location |  |  |
| Middle cerebral artery | 2.22 (1.05–4.90) | 0.05            | 1.00 (0.39–2.43) | 0.969           |
| Anterior cerebral artery | 0.96 (0.13–7.08) | 0.96           |                  |                 |
| Posterior cerebral artery | 1.66 (0.63–4.39) | 0.303           |                  |                 |
| Cortical involvement | 2.15 (1.21–3.84) | 0.008           | 2.00 (1.10–3.63) | 0.022           |
| Early ASM treatment | 12.72 (7.55–21.43) | <0.001          | 10.47 (5.63–19.47) | <0.001          |
| Early EEG findings |  |  |
| Abnormal | 2.36 (1.18–4.68) | 0.014           | 2.09 (0.90–4.82) | 0.084           |
| Diffuse slow | 2.96 (1.61–5.44) | <0.001          | 1.26 (0.65–2.43) | 0.487           |
| Focal slow | 1.79 (1.05–3.05) | 0.03            | 0.56 (0.30–1.32) | 0.110           |
| Interictal epileptiform discharges | 0.87 (0.21–3.60) | 0.855           |                  |                 |
| Laterialized periodic discharges | 1.76 (0.63–4.95) | 0.278           |                  |                 |
| Early seizure on EEG | 11.11 (4.42–27.93) | <0.001          | 2.01 (0.72–5.58) | 0.177           |
| Status epilepticus on EEG | 2.62 (0.36–19.00) | 0.34           |                  |                 |

N = 673.
aHR = adjusted hazard ratio; ASM = anti-seizure medication; CI = confidence interval; EEG = electroencephalogram; HR = hazard ratio; NIHSS = National Institutes of Health Stroke Scale.
Separate analyses were performed using the initial and the matched samples.

Third, we performed additional analyses:

a. We used causal mediation analysis to evaluate the relationship between reperfusion treatment and secondary hemorrhage with the occurrence of acute symptomatic seizures or post-stroke epilepsy. As suggested by other authors, we hypothesized that the epileptogenic effect could be mediated by the occurrence of secondary hemorrhage and not by the reperfusion treatment itself.

We performed this analysis in the Swiss (1) and Italian cohorts (n = 737) that had information on secondary hemorrhage. To control for exposure-outcome confounding, we applied the same matching for treatment propensity as described above and evaluated the average causal mediation effect (ie, ACME) of secondary hemorrhage on the outcome after matching. Confidence intervals (CIs) around mediation effect estimates were generated using bootstrapping with 1,000 simulations.

b. We evaluated the association between time to IV thrombolysis (door-to-needle) or successful recanalization after IA or mechanical recanalization and risks of acute symptomatic seizures or post-stroke epilepsy using linear and Cox regression analyses. For this, we only included the Swiss (2) cohort, given that data on time to IV thrombolysis (n = 415) or successful recanalization after interventional reperfusion treatment (n = 195) were collected in a standardized manner in this center.

We calculated the power to reject the null hypothesis (reperfusion treatment is associated with acute symptomatic seizures or post-stroke epilepsy) with a 5% type I error rate and using the risks and sample sizes obtained in the above calculations.

We followed established recommendations (ie, STROBE checklist). Analyses were performed using R version 1.4.53, using the packages “survival,” “survminer,” “MatchIt,” “stddiff,” “mediation,” and “ggplot2.”

**Results**

The overall registry included 4,229 individuals (mean age = 71 years old; 57% men, see Supplementary Table S1) from 8 European cohorts (Austria n = 459, Germany [1] n = 182, Germany [2] n = 311, Italy n = 399, Portugal n = 152, Spain = 511, Switzerland [1] n = 1,200, and Switzerland [2] n = 1,016). Four of these cohorts (Switzerland [1], Austria, Germany [2], and Italy) were part of the original SeLECT study and 2 of the 4 additional cohorts were previously published.1,36–39

Overall, 1,225 patients (29%) received reperfusion treatment, 196 (5%) experienced acute symptomatic seizures, and 232 (6%) had post-stroke epilepsy during follow-up of a median of 1.6 years (interquartile range [IQR] = 1.0–3.3).

The frequency of acute symptomatic seizures and post-stroke epilepsy among reperfusion groups were as follows: those who received reperfusion treatment, 6% and 8%; IV thrombolysis, 6% and 8%; IA thrombolysis 7% and 5%; and mechanical thrombectomy 8% and 5%, respectively. Median time to post-stroke epilepsy was 1.87 years (IQR = 1.0–3.2) in participants without reperfusion treatment and 1.1 years (IQR = 1.0–2.2) in participants with reperfusion treatment.

**Factors Associated with Acute Symptomatic Seizures or Post-Stroke Epilepsy**

Variables independently associated with a higher risk of acute symptomatic seizures in the overall cohort (n = 4,299; see Table 1) were NIHSS at admission ≥11 points (adjusted odds ratio [aOR] = 1.7, 95% CI = 1.2–2.3, p < 0.001), stroke located in the posterior cerebral artery (PCA) territory (aOR = 1.5, 95% CI = 1.1–2.0, p = 0.02), and stroke caused by large-artery atherosclerosis (aOR = 1.4, 95% CI = 1.0–2.0, p = 0.03). Stroke caused by small-vessel occlusion had a lower risk of acute symptomatic seizures (aOR = 0.2, 95% CI = 0.1–0.4, p < 0.001).

Factors independently associated with a shorter time to develop post-stroke epilepsy (see Table 2) included acute symptomatic seizures (adjusted hazard ratio [aHR] = 4.2, 95% CI = 2.9–6.0, p < 0.001), NIHSS at admission ≥11 points (aHR = 2.3, 95% CI = 1.7–3.2, p < 0.001), stroke involving the cerebral cortex (aHR = 2.1, 95% CI = 1.5–2.8, p < 0.001), and stroke caused by large-artery atherosclerosis (aHR = 1.5, 95% CI = 1.1–2.0, p = 0.004).

Reperfusion treatment was associated with a higher risk of post-stroke epilepsy (HR = 1.9, 95% CI = 1.4–2.4, p < 0.001; Fig 2A) and acute symptomatic seizures (OR = 1.5, 95% CI = 1.1–2.0, p = 0.008; Fig 3A) in univariable analysis. This association was nonsignificant after adjusting for other covariates in multivariable analysis (acute symptomatic seizures: aOR = 1.0, 95% CI = 0.8–1.4, p = 0.82; post-stroke epilepsy: aHR = 1.1, 95% CI = 0.8–1.4, p = 0.74).

**Predictive Role of Early EEG Findings for Post-Stroke Epilepsy**

We studied the role of EEG data within the 7 days of stroke onset for prediction of post-stroke epilepsy in a subset of 673 subjects from 3 cohorts (Portugal and
Hazard risk of time to post stroke epilepsy, before and after matching

A Reperfusion vs. no reperfusion

B IV thrombolysis vs. no reperfusion

C Mechanical thrombectomy vs. no reperfusion

Switzerland [1] and [2]) with available EEG data. Abnormal early EEG (HR = 2.36, 95% CI = 1.18–4.68), diffuse background slowing (HR = 2.96, 95% CI = 1.61–5.44), focal slowing (HR = 1.79, 95% CI = 1.05–3.05), and early seizures on EEG (HR = 11.11, 95% CI = 4.42–27.93) were significant predictors of post-stroke epilepsy in univariable Cox regression analysis (see Table 3A). None of the studied EEG abnormalities were
significant after multivariable correction for covariates (see Table 3B).

Reperfusion (IV, IA, or Mechanical Thrombectomy) Versus No Reperfusion Treatment
We matched 936 patients receiving reperfusion treatment and the same number of controls for reperfusion treatment propensity. No significant differences in baseline characteristics were observed after matching. Standardized differences (see Supplementary Table S1) and density plots (Fig 4A) showed improved balance for all included covariates.

After matching, we did not observe an association between reperfusion treatment and time to post-stroke epilepsy (HR = 1.05, 95% CI = 0.75–1.48, p = 0.74; see Fig 2A) and risk of acute symptomatic seizures (OR = 1.04, 95% CI = 0.70–1.55, p = 0.84; see Fig 3A).

Our study had an estimated power of 80% to reject the null hypothesis that reperfusion treatment is associated with acute symptomatic seizures and a 96% power to reject the null hypothesis that reperfusion treatment is associated with post-stroke epilepsy.

IV Thrombolysis Versus No Reperfusion Treatment
Before matching, participants who received IV thrombolysis had a shorter time to post-stroke epilepsy (HR = 1.98, 95% CI = 1.49–2.64, p < 0.001; see Fig 2B1) and higher risk of acute symptomatic seizures (OR = 1.52, 95% CI = 1.09–2.08, p = 0.001; see Fig 3B) compared to those that did not receive reperfusion treatment.

We matched 824 cases and controls for IV thrombolysis treatment propensity. No significant differences in baseline characteristics were observed after matching. Standardized differences (see Supplementary Table S2) and density plots (see Fig 4B) showed improved balance for all included covariates.

After matching, we did not observe any association between IV thrombolysis and time to post-stroke epilepsy (HR = 1.15, 95% CI = 0.80–1.65, p = 0.43; see Fig 2B2) or risk of acute symptomatic seizures (OR = 1.09, 95% CI = 0.72–1.62, p = 0.68; see Fig 3B).

Mechanical Thrombectomy Only Versus No Reperfusion
The risk of acute symptomatic seizures for individuals that received mechanical thrombectomy after ischemic stroke

Balance plots of propensity score matching, before and after matching
A Reperfusion vs. no reperfusion

B IV thrombolysis vs. no reperfusion

C Mechanical thrombectomy vs. no reperfusion

FIGURE 3: Risk of acute symptomatic seizures, before and after matching. Forest plot showing odd ratios and 95% confidence intervals (horizontal lines) for the risk of acute symptomatic seizures after acute ischemic stroke for each of the scenarios.

FIGURE 4: Density plots of propensity score matching. Density plots showing the covariate balance before and after matching, for each of the scenarios.
versus those that received no reperfusion treatment was significant (OR = 2.11, 95% CI = 1.28–3.31, \( p = 0.001 \); see Fig 3C), whereas the risk of post-stroke epilepsy was not (HR = 1.3, 95% CI = 0.74–2.32, \( p = 0.34 \); see Fig 2C1). After matching on the propensity to receive mechanical thrombectomy, we found no significant association with acute symptomatic seizures (OR = 1.24, 95% CI = 0.64–2.44, \( p = 0.51 \); see Fig 3C) or post-stroke epilepsy (HR = 1.4, 95% CI = 0.56–3.79, \( p = 0.42 \); see Fig 2C2). Standardized differences and density plots (see Fig 4C) showed improved balance for all included covariates.

**The Impact of Secondary Hemorrhage**

The effect of reperfusion treatments (IV, IA, and/or mechanical thrombectomy) on the likelihood of acute symptomatic seizures and post-stroke epilepsy was not mediated by secondary hemorrhage after matching for the propensity to receive reperfusion treatment. For acute symptomatic seizures, the unstandardized indirect effect was 0.99 (95% CI = 0.98–1.01, \( p = 0.91 \)). For post-stroke epilepsy, the unstandardized indirect effect was 0.98 (95% CI = 0.94–1.01, \( p = 0.26 \)). The results were comparable when comparing IV thrombolysis with no reperfusion treatment (acute symptomatic seizure: unstandardized indirect effect 1.00, 95% CI = 0.98–1.01, \( p = 0.90 \); and post-stroke epilepsy: 0.98, 95% CI = 0.93–1.01, \( p = 0.26 \)). Further sensitivity analyses were not performed as the indirect effects were not significant.\(^{40}\)

**The Impact of Time to Thrombolysis and Successful Recanalization**

Neither time to thrombolysis nor successful recanalization after IA or mechanical thrombolysis were associated with the risk of acute symptomatic seizures (\( p = 0.18 \) and \( p = 0.30 \), respectively) or post-stroke epilepsy (\( p = 0.59 \) and \( p = 0.53 \)) in univariable and multivariable analyses in the Swiss (2) cohort.

**Discussion**

We used data from a large multicenter registry to show that the risk of post-stroke seizures is associated with stroke severity, etiology, and location. Acute symptomatic seizures were the strongest predictor of post-stroke epilepsy. Findings on routine EEG within the first week were not independently associated with post-stroke epilepsy.

Reperfusion treatment, be it IV or IA thrombolysis or mechanical thrombectomy, for ischemic stroke was not associated with an increased risk of post-stroke seizures after correction for treatment selection bias by PSM. Our analyses suggest that the previously described associations were potentially confounded by treatment selection, with reperfusion therapy being more likely to occur in those with more severe cerebral infarcts.

The risk factors for acute symptomatic seizures and post-stroke epilepsy identified in the current investigation were comparable with previous studies.\(^{1,7}\) Acute symptomatic seizures were associated with more severe strokes (NIHSS at admission \( \geq 11 \) points) and stroke etiology (higher risk in large-artery atherosclerosis and lower risk in small-vessel occlusion). Similarly, the risk of post-stroke epilepsy was highest after more severe strokes (NIHSS at admission \( \geq 11 \) points) and strokes caused by large-artery atherosclerosis. The most relevant predictor of post-stroke epilepsy were acute symptomatic seizures. Acute symptomatic seizures indicate an increased susceptibility to generate seizures following an insult thus being related to a markedly increased risk of developing epilepsy. Moreover, seizure-related head injuries, in particular head trauma, may increase the risk of subsequent seizures and post-traumatic epilepsy.\(^{41}\)

Infarct location was identified as a potential risk factor. Acute symptomatic seizures were associated with PCA territory infarcts. Future studies refining the areas affected with lesion-symptom mapping could address these aspects and clarify whether there are particular areas that are most likely to be associated with developing seizures (eg, the posterior medial temporal lobe). For post-stroke epilepsy, however, the most relevant prognostic factor was cortical damage, whereas the arterial location of the infarct did not play a significant role. This confirms the clinical observation that lesions in all supratentorial vascular territories may cause post-stroke epilepsy, particularly if the cerebral cortex is involved.

In a subset of our cohort with EEG within the first 7 days following stroke, EEG findings did not predict post-stroke epilepsy. This is in contrast with 2 smaller previous studies.\(^{12,42}\) The differences between studies may be explained by adjustment for covariates. In our study, we rigorously adjusted our models for potential covariates, including the presence of acute symptomatic seizures. The EEG findings that were significant in univariable analysis were not relevant after correction for clinical variables and stroke location. In contrast, one previous study used a prospective assessment of repeated EEG recordings using a 64-channel system, which may have increased the sensitivity and specificity of the EEG findings.\(^{12}\) Thus, the chosen variables which were extracted from the visual analysis of early routine EEG may not be reliable independent biomarkers for the development of post-stroke epilepsy, whereas extended EEG monitoring with a 64-channel system showed some benefit in a previous study.\(^{12}\)
Acute symptomatic seizures after stroke are uncommon but they have a negative impact on outcome. Seizures and discharges on EEG are associated with increased metabolic stress and the release of extracellular glutamate, which may contribute to neurotoxicity after stroke. Thus, reports that reperfusion treatment after stroke may cause seizures raised concern.

Proposed potential mechanisms are a pro-epileptic effect of a maintained penumbra or blood–brain barrier disruption following reperfusion or a direct proconvulsive effect of rtPA. Treatment using rtPA has also been found to increase the likelihood of hemorrhagic transformation, which in turn may increase the risk of post-stroke seizures.

The optimal study design to assess the contribution of reperfusion treatment to seizure risk after stroke is a randomized controlled trial, but such a trial would require a large sample size. It would also be unethical to withhold reperfusion treatment due to its proven benefits. Thus, the best available approach is to analyze prospectively acquired data from nonrandomized cohort studies. Such comparisons are confounded by treatment selection. We used PSM, to reduce treatment selection bias, to mimic important aspects of a randomized controlled trial.

Our results show that the previously reported association of reperfusion treatment with seizures can be attributed to treatment selection bias. Both time to post-stroke epilepsy and the occurrence of acute symptomatic seizures were associated with higher stroke severity, infarcts involving the cortex and large-artery atherosclerotic etiology. These factors are also associated with a higher propensity to receive reperfusion treatment (see Supplementary Table S1). This observation highlights the importance of adequately addressing confounders in nonrandomized trials and the need for a critical approach when interpreting data from nonrandomized studies.

We found no differences in the risk of acute or post-stroke epilepsy after correction for reperfusion treatment propensity, neither in the overall population nor in the subgroup of individuals with stroke affecting the cortex in the MCA territory. As it has been previously hypothesized that rtPA may have a direct pro-seizure effect, we compared participants who received IV thrombolysis to those that had no reperfusion treatment. We did not find an effect of IV thrombolysis on post-stroke seizures in the matched cohorts. It has also been hypothesized that mechanical thrombectomy may have pro-convulsive effects. In our study, there was no effect of mechanical thrombectomy on post-stroke seizures in the matched comparison.

Thrombolysis and thrombectomy reduce overall disability after stroke. These treatments did, however, not appear to reduce the risk of post-stroke seizures and epilepsy. We speculate that, despite the removal of the thrombus, there may be ongoing excitotoxicity, inflammation, or blood–brain-barrier disruption leading to the development of an epileptogenic network. This highlights the importance to develop other strategies to prevent seizures and epileptogenesis after stroke.

Secondary hemorrhage has been described as a risk factor for seizures after stroke. Our results show that secondary hemorrhage after reperfusion treatment or IV thrombolysis does not mediate the likelihood of having post-stroke seizures. This result is in line with the lack of association between reperfusion treatment and occurrence of seizures.

Time to thrombolysis and successful recanalization have been shown to be strongly associated with function outcome and reduced mortality. This is the first study examining the effect of these variables on the risk of post stroke epilepsy. Door-to-needle times and recanalization rates were not associated with the risk of acute symptomatic seizures or post-stroke epilepsy.

Our study has several strengths. First, the assembled cohort is one of the largest studies of ischemic stroke participants with long follow-up regarding seizures. We had a high power to detect potential differences between patients with or without reperfusion treatment. Second, our matching approach mimics some of the aspects of a randomized controlled trial and reduces confounds. Third, using data from 8 international centers provides support for the generalizability of the findings. Fourth, our results are applicable to a wide range of patients with ischemic stroke and also in those with stroke affecting the cortex in the MCA territory, that may have a larger a priori risk of seizures. Our results were valid for all types of reperfusion treatment and also in the subgroups receiving IV thrombolysis or mechanical thrombectomy.

Our study has limitations. First, we only considered clinically apparent seizures and may have underestimated the incidence of nonconvulsive seizures. Using continuous EEG after stroke might have increased the detection of seizures with subtle or no clinical signs but it would not be feasible in a retrospective study not designed to robustly detect asymptomatic seizures. Second, data on time to IV thrombolysis and recanalization rates after IA or mechanical reperfusion treatment were collected in a standardized manner only in the Swiss (2) cohort, restricting these analyses to a smaller sample. Last, even though we reduced between-group baseline differences using PSM, we cannot rule out the possibility of residual confounding due to unmeasured or unknown factors.
To conclude, our results describe risk factors for seizures following stroke and show that reperfusion treatment after ischemic stroke is not associated with an increased risk of acute symptomatic seizures or of post-stroke epilepsy.

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Author Contributions
C.F.A. and M.G. contributed to the conception and design of the study. C.F.A., N.D., E.C., A.F., P.S., N.S., G.B., J.S., L.S., L.I., M.K., L.A., E.S., J.A.S., M.W., T.J.O., J.N.W., G.L.G., A.R., F.J., G.M., M.V., G.G., J.C., S.E., P.L., F.R., F.B., C.B., A.R.P., T.P.M., B.T., M.R.K., J.S.D., J.W.S., B.T., M.J.K., and M.G. contributed to the acquisition and analysis of the data. C.F.A. and M.G. contributed to drafting the text or preparing the figures.

Potential Conflicts of Interest
The authors declared no conflict of interest.

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