Genetic Polymorphism of Glucokinase on the Risk of Type 2 Diabetes and Impaired Glucose Regulation: Evidence Based on 298, 468 Subjects

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Abstract

Background: Glucokinase (GCK) is the key glucose phosphorylation enzyme which has attracted considerable attention as a candidate gene for type 2 diabetes (T2D) based on its enzyme function as the first rate-limiting step in the glycolysis pathway and regulates glucose-stimulated insulin secretion. In the past decade, the relationship between GCK and T2D has been reported in various ethnic groups. To derive a more precise estimation of the relationship and the effect of factors that might modify the risk, we performed this meta-analysis.

Methods: Databases including Pubmed, EMBASE, Web of Science and China National Knowledge Infrastructure (CNKI) were searched to find relevant studies. Odds ratios (ORs) with 95% confidence intervals (CIs) were used to assess the strength of association.

Results: A total of 24 articles involving 88,229 cases and 210,239 controls were included. An overall random-effects per-allele OR of 1.06 (95% CI: 1.03–1.09; P<10^-4) was found for the GCK –30G>A polymorphism. Significant results were also observed using dominant or recessive genetic models. In the subgroup analyses by ethnicity, significant results were found in Caucasians; whereas no significant associations were found among Asians. In addition, we found that the –30G>A polymorphism is a risk factor associated with increased impaired glucose regulation susceptibility. Besides, –30G>A homozygous was found to be significantly associated with increased fasting plasma glucose level with weighted mean difference (WMD) of 0.15 (95%: 0.05–0.24, P=0.001) compared with G/G genotype.

Conclusions: This meta-analysis demonstrated that the –30G>A polymorphism of GCK is a risk factor associated with increased T2D susceptibility, but these associations vary in different ethnic populations.

Introduction

Type 2 diabetes (T2D) is a complex metabolic disease characterised by hyperglycaemia, insulin resistance, impaired insulin secretion due to pancreatic β-cell defects and increased hepatic glucose production. Despite much investigation, the causes underlying the development and progression of T2D have not been fully elucidated, accumulated evidence suggests that multiple genetic and environmental factors, as well as the interplay between these factors, determine the phenotype. Although the genetic contribution to T2D is well recognized, the current set of 56 established susceptibility loci, identified primarily through large-scale genome-wide association studies (GWAS), captures at best 10% of familial aggregation of the disease [1–3]. This has maintained interest in other biochemical and genetic factors that might contribute to the underlying pathophysiology of the disease.

Glucokinase (GCK) is the key glucose phosphorylation enzyme responsible for the first rate-limiting step in the glycolysis pathway and regulates glucose-stimulated insulin secretion from pancreatic β-cells and glucose metabolism in the liver [4]. Inactivating GCK mutations lead to maturity-onset diabetes of the young and neonatal diabetes [5–7], whereas activating GCK mutations cause persistent hyperinsulinaemic hypoglycaemia [8–11]. Moreover, a common variant (–30G>A), rs1799884 in the pancreatic beta cell-specific promoter of GCK has been shown to be associated with increased risk of type 2 diabetes, hyperglycaemia and impaired beta cell function [12–16]. Furthermore, GCK –30 A>G has
been conclusively associated with fasting glucose in European populations [17].

To date, many case–control studies have been carried out to investigate the role of the GCK −30G>A polymorphism in the development of T2D among various populations. Genetic association studies can be problematic to reproduce due to insufficient power, multiple hypothesis testing, population stratification, source of controls, publication bias, and phenotypic heterogeneity. In addition, with the increased studies in recent years among Asian, and other populations, there is a need to reconcile these data. We therefore conducted a comprehensive meta-analysis to quantify the overall risk of GCK −30G>A polymorphism on developing T2D.

**Materials and Methods**

**Literature Search Strategy**

Genetic association studies published before the end of Sep. 2012 on T2D and polymorphisms within GCK gene were identified through a search of PubMed, ISI Web of Science, EMBASE and CNKI (Chinese National Knowledge Infrastructure) without language restrictions. Search term combinations were keywords relating to the glucokinase gene (e.g., “glucokinase”, “GCK”, and “MODY 2”) in combination with words related to T2D (e.g., “type 2 diabetes mellitus”, “T2DM”, “type 2 diabetes”, “T2D”, “non-insulin-dependent diabetes mellitus” and “NIDDM”) and polymorphism or variation. The search was supplemented by reviews of reference lists for all relevant studies and review articles. The major inclusion criteria were (a) original papers containing independent data, (b) case-control or cohort studies and (c) genotype distribution information or odds ratio (OR) with its 95% confidence interval (CI) and P-value. The major reasons for exclusion of studies were (a) overlapping data and (b) case-only studies, family-based studies and review articles.

**Data Extraction**

Data extraction was performed independently by two reviewers, and differences were resolved by further discussion among all authors. For each included study, the following information was extracted from each report according to a fixed protocol: first author’s surname, publication year, definition and numbers of cases and controls, diagnostic criterion, frequency of genotypes, source of controls, age, gender, body mass index (BMI), Hardy–Weinberg equilibrium status, ethnicity and genotyping method. Not all researchers use the same SNP, we report herein 2 common SNPs (rs1799884 and rs4607517), as these SNPs are in complete disequilibrium ($r^2 = 1$) [18].

**Statistical Methods**

The strength of association between −30G>A polymorphism of GCK and T2D risk was assessed by odds ratio (OR) with the corresponding 95% confidence interval (CI). We first used the chi square test to check if there was significant deviation from Hardy–Weinberg equilibrium among the control subjects in each study. The meta-analysis examined the association between each polymorphism and the risk of T2D for the: (i) allele contrast, (ii) dominant, and (iii) recessive model. For continuity variable, weighted mean difference (WMD) was used to pool results from studies.

Heterogeneity across individual studies was calculated using the Cochran chi-square Q test followed by subsidiary analysis or by random-effects regression models with restricted maximum likelihood estimation [19–21]. Random-effects and fixed-effect summary measures were calculated as inverse variance-weighted average of the log OR. The results of random-effects summary were reported in the text because it takes into account the variation between studies. In addition, we investigated potential sources of identified heterogeneity among studies by stratification according to the number of T2D cases (≥1000 and <1000), ethnic group and diagnostic criteria (WHO, ADA or other criterion). Ethnic group was defined as Asians, Caucasians (i.e. people of European origin) and others (e.g. Tunisian and African-American). The Z test was used to determine the significance of the pooled OR. Gender distribution in cases and controls, genotyping method and mean age of cases and controls were analysed as covariates in meta-regression. The transcript expression level trends by genotypes were evaluated by using general linear model.

We assessed publication bias by using an ancillary procedure attributed to Egger et al. [22], which uses a linear regression approach to measure funnel plot asymmetry on the natural logarithm of the OR. Sensitivity analysis was performed by removing each individual study in turn from the total and re-analysing the remainder. This procedure was used to ensure that no individual study was entirely responsible for the combined results. All statistical analyses were carried out with the Stata software version 10.0 (Stata Corporation, College Station, TX, USA). The type I error rate was set at 0.05. All the P-values were for two-sided analysis.

**Results**

**Characteristics of Studies**

Study selection process was shown in Figure S1. In all, we included 36 data points from 24 studies in this meta-analysis, with a total of 88,229 cases and 210,239 controls [12–14,16,18,23–41]. The distribution of genotypes in the controls was consistent with Hardy–Weinberg equilibrium in all studies. Of the cases, 75% were Caucasians 21% were Asians and 4% were of other ethnic origins. The main study characteristics were summarized in Table 1.

**Association of GCK −30G>A Variant with T2D**

For T2D risk and the −30G>A polymorphism of GCK, our meta-analysis gave an overall OR of 1.06 (95% CI: 1.03–1.09; $P<10^{-4}$; Fig. 1). Significantly increased T2D risks were also found under dominant (OR = 1.08; 95% CI: 1.01–1.19; $P = 0.003$) and recessive genetic models (OR = 1.12; 95% CI: 1.01–1.25; $P = 0.008$). This analysis is based on pooling of data from a number of different ethnic populations. When stratifying for ethnicity, an OR of 1.08 (95% CI: 1.04–1.12; $P = 10^{-4}$), 1.01 (95% CI: 0.98–1.05; $P = 0.53$) and 1.13 (95% CI: 1.04–1.24; $P = 0.006$) resulted for the A allele, among Caucasian, Asian and other ethnic populations, respectively. Similar results were also detected using dominant and recessive genetic models (Table 2). When studies were stratified for sample size, significant risks were found among studies with small sample size in all genetic model (A allele: OR = 1.14, 95% CI: 1.04–1.25; dominant model: OR = 1.17, 95% CI: 1.06–1.28; recessive model: OR = 1.23, 95% CI: 1.05–1.43). Positive results still maintained for large sample size studies in all genetic models. Subsidiary analyses of diagnostic criterion yielded a per-allele OR for WHO criterion of 1.06 (95% CI: 1.01–1.11; $P = 0.02$), ADA criterion of 1.13 (95% CI: 0.99–1.28; $P = 0.06$) and for other criterion of 1.05 (95% CI: 0.99–1.11; $P = 0.12$).

Significant heterogeneity was present among the included studies ($P<0.05$). In meta-regression analysis, mean age of cases ($P = 0.31$) and controls ($P = 0.24$) and genotyping method ($P = 0.96$) did not significantly explain such heterogeneity. By
Table 1. Characteristics of the studies included in the meta-analysis.

| Reference | Year | Ethnicity | Case | Control | No. of case | No. of control | Genotyping method |
|-----------|------|-----------|------|---------|-------------|----------------|-------------------|
| Cauchi [26] | 2012 | Arab | T2D per WHO criteria | Healthy | 2639 | 1997 | TaqMan |
| Cho [27] | 2011 | Asian | T2D patients | Non-diabetic participants | 6952 | 11865 | Genechip |
| Kooner [28] | 2011 | Asian | T2D patients | Non-diabetic participants | 5561 | 14458 | Genechip |
| Hu [29] | 2010 | Chinese | T2D per WHO criteria | Normal glucose tolerance | 3410 | 3412 | MassArray |
| Murad [30] | 2010 | British | T2D patients | Non-diabetic participants | 1551 | 2993 | TaqMan |
| Tam [31] | 2010 | Chinese | T2D per WHO criteria | Normal fasting glucose | 1320 | 1595 | MassArray |
| Dupuis [32] | 2010 | European, American, Austrian | T2D per WHO/ADA criteria | Non-diabetic participants | 40655 | 87022 | SNPstream, Genechip, TaqMan, MassArray |
| Ezzidi [33] | 2009 | Tunisian | T2D per ADA criteria | Normoglycemic participants | 865 | 505 | TaqMan |
| Prokopenko [18] | 2009 | European | T2D per WHO criteria; T2D patients | Normal glucose tolerance; Non-diabetic participants | 11785 | 49799 | Genechip |
| Reiling [34] | 2009 | Dutch | T2D per WHO criteria | Normal glucose tolerance | 2498 | 1912 | TaqMan |
| Qi [35] | 2009 | Chinese | T2D per WHO criteria | Normal fasting glucose | 416 | 1877 | SNPstream |
| Ma [36] | 2009 | Chinese | T2D per WHO criteria | Non-diabetic participants | 279 | 110 | RFLP |
| Cauchi [37] | 2008 | French | T2D per ADA criteria | Normoglycemic participants | 2637 | 4159 | TaqMan |
| Vaxillaire [12] | 2008 | French | T2D per ADA criteria | Normoglycemic participants | 292 | 2752 | TaqMan |
| Holmkööst [38] | 2008 | Swedish | T2D per ADA criteria | Non-diabetic participants | 1988 | 15019 | TaqMan |
| Winckler [39] | 2007 | British | T2D per WHO criteria | Non-diabetic participants | 2348 | 3561 | MassArray |
| Bonnaycastle [40] | 2006 | Finnish | T2D per WHO criteria | Normal glucose tolerance | 784 | 617 | MassArray |
| Rose [14] | 2005 | Dane | T2D per WHO criteria | Normal glucose tolerance | 1408 | 4773 | MassArray |
| März [13] | 2004 | Austrian | T2D per ADA criteria | Non-diabetic participants | 463 | 830 | RFLP |
| Rissinen [41] | 1998 | Finnish | T2D patients | Normal glucose tolerance | 36 | 294 | SSCP |
| Yamada [16] | 1997 | Japanese | T2D per WHO criteria | Normal glucose tolerance | 94 | 321 | RFLP |
| Löfl [42] | 1997 | Swedish | T2D per WHO criteria | Healthy | 31 | 158 | SSCP |
| Shimokawa [43] | 1994 | Japanese | T2D patients | Non-diabetic participants | 240 | 111 | SSCP |
| Chiu [44] | 1994 | American Blacks | T2D per NDDG criteria | Non-diabetic participants | 77 | 99 | SSCP |
contrast, ethnicity ($P = 0.02$) and the sample size in cases ($P = 0.01$) was significantly correlated with the magnitude of the genetic effect, explaining 11% and 16% of the heterogeneity, respectively. Since significant between-study heterogeneity still maintained in Caucasian subgroup, hence studies on Caucasian populations is the main source of heterogeneity.

Association of GCK $-30G>A$ Variant with Impaired Glucose Regulation

To investigate how glucose metabolism was related to glucokinase, we analyzed individuals with impaired glucose regulation (impaired glucose tolerance and/or impaired fasting glucose). The data on genotypes of the $-30G>A$ polymorphism among...
impaired glucose regulation cases and controls were available in 5 studies (including 3177 cases and 8970 controls). In the overall analysis, the 230G>A polymorphism of GCK was significantly associated with elevated impaired glucose regulation risk with a per-allele OR of 1.23 [95% CI: 1.14–1.32; \( P(Z) = 0.25; P(Q) = 0.73 \); Fig. 2]. Significant associations were also found under dominant [OR = 1.24; 95% CI: 1.13–1.35; \( P(Z) < 10^{-5} \); \( P(Q) = 0.36 \)] and recessive [OR = 1.52; 95% CI: 1.20–1.91; \( P(Z) = 0.004; P(Q) = 0.77 \)] genetic model.

### Table 2. Meta-analysis of the GCK –30G>A polymorphism on type 2 diabetes risk.

| Sub-group analysis | No. of data sets | No. of case/control | A allele | Dominant model | Recessive model |
|-------------------|------------------|---------------------|----------|----------------|----------------|
|                   |                  |                     |          | OR (95%CI)  | P(Z) | P(Q)* | OR (95%CI)  | P(Z) | P(Q)* | OR (95%CI)  | P(Z) | P(Q)* |
| Overall           | 36               | 88229/210239        | 1.06 (1.03–1.09) | \(< 10^{-4}\) | 0.003 | 1.08 (1.01–1.19) | 0.003 | 0.009 | 1.12 (1.01–1.25) | 0.008 | 0.001 |
| Ethnicity         |                  |                     |          | \(P(Z)\) | \(P(Q)\) | \(P(Z)\) | \(P(Q)\) | \(P(Z)\) | \(P(Q)\) |
| Caucasian         | 23               | 66376/173889        | 1.08 (1.04–1.12) | \(< 10^{-4}\) | 0.002 | 1.13 (1.04–1.21) | 0.0007 | 0.006 | 1.17 (1.06–1.29) | 0.002 | \(< 10^{-4}\) |
| Asian             | 9                | 18272/33749         | 1.01 (0.98–1.05) | 0.53 | 0.77 | 0.96 (0.84–1.10) | 0.57 | 0.82 | 1.05 (0.93–1.18) | 0.76 | 0.34 |
| Others            | 4                | 3581/2601           | 1.13 (1.04–1.24) | 0.006 | 0.49 | 1.06 (1.03–1.13) | 0.008 | 0.63 | 1.12 (1.05–1.24) | 0.03 | 0.52 |
| Sample size       |                  |                     |          | \(P(Z)\) | \(P(Q)\) | \(P(Z)\) | \(P(Q)\) | \(P(Z)\) | \(P(Q)\) |
| Small             | 0.001            | 15                 | 4640/12804 | 1.14 (1.04–1.25) | 0.006 | 0.09 | 1.17 (1.06–1.28) | 0.009 | 0.21 | 1.23 (1.05–1.43) | 0.01 | 0.10 |
| Large             | 21               | 83589/197435        | 1.04 (1.01–1.07) | 0.004 | 0.03 | 1.09 (1.02–1.23) | 0.006 | 0.13 | 1.11 (1.07–1.35) | 0.002 | 0.18 |
| Diagnostic criterion |                  |                     |          | \(P(Z)\) | \(P(Q)\) | \(P(Z)\) | \(P(Q)\) | \(P(Z)\) | \(P(Q)\) |
| WHO criterion     | 19               | 72385/169960        | 1.06 (1.01–1.11) | 0.02 | 0.12 | 1.07 (1.02–1.19) | 0.01 | 0.31 | 1.09 (1.05–1.17) | 0.003 | 0.29 |
| ADA criterion     | 7                | 6247/23265          | 1.13 (0.99–1.28) | 0.06 | \(< 10^{-4}\) | 1.06 (0.97–1.16) | 0.13 | \(< 10^{-4}\) | 1.18 (0.94–1.47) | 0.20 | \(< 10^{-5}\) |
| Other criterion   | 10               | 9597/17014          | 1.05 (0.99–1.11) | 0.12 | 0.24 | 1.07 (0.97–1.19) | 0.16 | 0.35 | 1.02 (0.87–1.19) | 0.76 | 0.09 |

*Cochran’s chi-square Q statistic test used to assess the heterogeneity in subgroups.

*Cochrane’s chi-square Q statistic test used to assess the heterogeneity between subgroups.

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**Figure 2. Forest plot for the overall association between the GCK –30G>A polymorphism and impaired glucose regulation risk.**

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impaired glucose regulation cases and controls were available in 5 studies (including 3177 cases and 8970 controls) and tested in the overall analysis, the –30G>A polymorphism of GCK was significantly associated with elevated impaired glucose regulation risk with a per-allele OR of 1.23 [95% CI: 1.14–1.32; \( P(Z) < 10^{-5} \); \( P(Q) = 0.73 \); Fig. 2]. Significant associations were also found under dominant [OR = 1.24; 95% CI: 1.13–1.35; \( P(Z) < 10^{-5} \); \( P(Q) = 0.36 \)] and recessive [OR = 1.52; 95% CI: 1.20–1.91; \( P(Z) = 0.004; P(Q) = 0.77 \)] genetic model.

### Association of GCK –30G>A Variant with Fasting Plasma Glucose Level

The data on fasting plasma glucose (FPG) level among subjects stratified by genotype of –30G>A polymorphism were available.
in 6 studies, including 31,616 subjects. Significant increases of fasting plasma glucose level were observed in A allele carriers compared with non-carriers. Using the random-effects model, compared with G/G genotype, the WMD for heterozygous were 0.07 [95%: 0.05–0.08, \(P(Z) < 10^{-5}\), \(P(Q)< 10^{-5}\), Figure 3] and homozygous 0.15 [95%:0.05–0.24, \(P(Z) = 0.001\), \(P(Q)<10^{-5}\)], respectively.

Sensitivity Analyses and Publication Bias

Sensitivity analysis indicated that no single study influenced the pooled OR qualitatively, suggesting that the results of this meta-analysis are stable (data not shown). The shape of the funnel plots was symmetrical (Figure S2). The statistical results still did not show publication bias in these studies (Begg test, \(P = 0.21\); Egger test, \(P = 0.60\)).

Discussion

Large sample and unbiased epidemiological studies of predisposition genes polymorphisms could provide insight into the in vivo relationship between candidate genes and diseases. This is the most comprehensive meta-analysis examining the GCK 230G>A polymorphism and the relationship to T2D risk. Its strength was based on the accumulation of published data giving greater information to detect significant differences. In total, the present meta-analysis combined 24 studies for T2D including 88,229 cases and 210,239 controls. Our results demonstrated that a modest association existed between the 230G>A variant of GCK and T2D risk.

In meta-analysis, heterogeneity evaluation was always conducted in statistical analysis. Thus, several subgroup meta-analyses were performed. In the stratified analysis by ethnicity, significant associations were found in Caucasians for the polymorphism in all genetic models; while no associations were detected among Asians. There are several possible reasons for such differences. First, the distribution of the A allele varies extensively between different races, with a prevalence of ~23% among Asians and ~17% among Caucasians. Therefore, additional studies are warranted to further validate ethnic difference in the effect of this polymorphism on T2D risk. In addition, different populations usually have different linkage disequilibrium patterns. A polymorphism may be in close linkage with another nearby causal variant in one ethnic population but not in another. GCK 230G>A polymorphism may be in close linkage with different nearby causal variants in different populations. Moreover, clinical heterogeneity like age, sex ratio, dietary, years from onset and disease severity may also explain the discrepancy. Finally, such different results could also be explained by study design or sample size. As significant between-study heterogeneity was found among Caucasian subgroup, so the result must be interpreted with caution since the Caucasian population reports in the subgroup analysis include a mixture of populations from very distant countries.

The present study also provides evidence that the 230G>A polymorphism of GCK influences susceptibility to phenotypes of impaired glucose regulation. In terms of genetic versus environmental influences on T2D susceptibility, this finding supports previous heritability studies, including a Danish twin study, generating considerably higher heritability estimates for the impaired glucose tolerance state compared with manifest T2D [42]. We also observed a significant effect of the 230G>A variant on FPG levels. This confirms a recent observation in a French study.

Figure 3. Meta-analysis of weighted mean differences (WMD) of fasting plasma glucose levels between GG and GA genotype of 230G>A polymorphism.

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Therefore, this variant may have a non-negligible impact on human health. Indeed, there is strong evidence that even small changes in FPG, well below the impaired fasting glucose threshold of 6.1 mmol/L, may be associated with risk of cardiovascular morbidity and mortality [43,44]. In this context, the GCK (−30G) allele was previously shown to be associated with type 2 diabetes and increased risk for coronary heart diseases in both diabetic and nondiabetic samples [13]. The exact mechanism by which the GCK −30A allele causes hyperglycemia is uncertain, but its effect seems constant throughout the lifespan, although insulin secretion is known to decrease with age in the general population. This is in accordance with the constant effect of the GCK −30A allele on fasting glucose reported in several groups of normoglycemic subjects whose median age varied from 8 to 72 years [12].

A number of factors predict T2D; however, detailed pathogenesis mechanisms of T2D remain a matter of speculation. GCK is a key regulatory enzyme in the pancreatic β-cell, and it plays a crucial role in determining the threshold for glucose-stimulated insulin secretion. Heterozygous inactivating mutations in GCK cause maturity-onset diabetes of the young subtype 2, in which hyperglycemia is present from birth. The decreased expression of functional GCK seems to be the cause of the observed hyperglycemia among maturity-onset diabetes of the young subtype 2 patients. The mechanism by which the −30G>A polymorphism causes hyperglycemia is uncertain. Previously published studies suggested that the minor A allele or a genetic variant with which it is in linkage disequilibrium may alter the expression of GCK [14].

In interpreting the results, some limitations of this meta-analysis should be addressed. Firstly, in the subgroup analyses, different ethnicities were pooled in the other ethnic group which may bring in some heterogeneity. As studies among the other ethnic group are currently limited, further studies including a wider spectrum of subjects should be carried to investigate the role of these variants in different populations. Secondly, our results were based on unadjusted estimates, while a more precise analysis should be conducted if all individual raw data were available, which would allow for the adjustment by other co-variants including age, drinking status, obesity, cigarette consumption, and other lifestyle. Thirdly, only published studies were included in this meta-analysis. Therefore, publication bias may have occurred, even though the use of a statistical test did not show it.

Despite these limitations, this meta-analysis suggests that GCK −30G>A polymorphism was significantly associated with increased risk of T2D, particularly in Caucasian population. In addition, GCK −30A allele is a true risk factor for the development of impaired glucose regulation, having a significant impact on FPG level.

Supporting Information

Figure S1 Study selection process. (TIF)

Figure S2 Beggs’s funnel plot of GCK −30G>A polymorphism and T2D risk. (TIF)

Checklist S1

Author Contributions

Conceived and designed the experiments: DF YH ZL. Performed the experiments: DF XC YM MC DL ZL. Analyzed the data: YM DL. Contributed reagents/materials/analysis tools: XC YM MC XY. Wrote the paper: DF XC YM YH ZL.

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