HALO ASSEMBLY BIAS IN HIERARCHICAL STRUCTURE FORMATION

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ABSTRACT

We investigate the origin of halo assembly bias, the dependence of halo clustering on assembly history. We relate halo assembly to peak properties measured in the Lagrangian space of the initial linear Gaussian random density field and show how these same Lagrangian properties determine large-scale bias. We focus on the two regimes where assembly bias has been observed to be significant: at masses very large and very small compared to the nonlinear mass scale. At high masses, we show that assembly bias is expected from the statistics of the peaks of Gaussian random fluctuations, and we show that the extent of assembly bias found in N-body simulations of rare halos is in excellent agreement with our theoretical prediction. At low masses, we argue that assembly bias largely arises from a subpopulation of low-mass halos whose mass accretion has ceased. Due to their arrested development, these halos naturally become unbiased, in contrast to their antibiased peers. We show that a simple toy model incorporating these effects can roughly reproduce the bias trends found in N-body simulations.

Subject headings: cosmology: theory — dark matter — large-scale structure of universe — methods: numerical

1. INTRODUCTION

Hierarchical structure formation is a basic prediction of the highly successful cold dark matter cosmological model, and a key element of hierarchical growth is the formation of bound virtualized objects called dark matter halos. Dark matter halos are not distributed uniformly throughout the universe, but rather compose a beaded filamentary network, sometimes referred to as the cosmic web (Bond et al. 1996). The clustering of the rare, massive dark matter halos is enhanced relative to the general mass distribution (Kaiser 1984; Efstathiou et al. 1988; Cole & Kaiser 1989; Bond et al. 1991; Mo & White 1996; Sheth & Tormen 1999), an effect known as bias. The more massive the halo, the larger the bias. As a result, the mass of halos hosting a given population of objects is sometimes inferred by measuring their degree of clustering—allowing a statistical route to the notoriously difficult problem of measuring masses of cosmological objects. There are many applications for statistical mass determination, including (for example) cluster mass-observable self-calibration (e.g., Lima & Hu 2004; Wu et al. 2008) or reconstruction of the halo occupation distribution of galaxies or quasars (e.g., Cooray & Sheth 2002).

Since halos of a given mass can differ in their formation histories and internal properties, a natural question arises: do these details affect halo clustering? Numerical simulations are now able to produce samples with sufficient statistics to test for the dependence of clustering on history and properties, and recent numerical simulations reveal subtle dependences of halo clustering on formation redshift, internal structure, and spin (Gao et al. 2005; Wechsler et al. 2006; Harker et al. 2006; Bett et al. 2007; Wetzel et al. 2007; Jing et al. 2007; Gao & White 2007; Angulo et al. 2008).

Using a marked correlation function, Sheth & Tormen (2004) found that close halo pairs tend to have earlier formation times than more distant pairs, work which was extended and confirmed by Harker et al. (2006). Gao et al. (2005) found that later forming, low-mass halos are less clustered than typical halos of the same mass at the present. Wechsler et al. (2006) and Wetzel et al. (2007) found a similar dependence on halo formation time (defined in a slightly different way), showing that the trend reversed for more massive halos and that the clustering depended on halo concentration. Recently, Gao & White (2007) showed that a wide variety of halo properties correlate with bias, including spin and substructure content, while Li et al. (2008) showed that different definitions of halo age, when applied to the same halo samples, give different amounts of assembly bias.

What remains unclear from this numerical work is the physical origin of this effect. Gao et al. (2005) argued that assembly bias is inconsistent with the Press & Schechter (1974) model, which explicitly predicts that halo abundance (and hence clustering) depends only on halo mass and no other internal properties related to assembly history. Attempts to explain assembly bias based on the Press-Schechter approach i.e., the excursion set using a sharp k-space filter would therefore appear doomed from the outset; several workers have explored modifications to the Press-Schechter model in the hopes of accounting for assembly bias. Sandvik et al. (2007) considered excursion set mass functions using barrier heights determined from the ellipsoidal collapse model of Bond & Myers (1996), but were unable to account for the magnitude of assembly bias seen in simulations. Mo et al. (2005) suggested that “preheating,” arising from the partial collapse of structure into sheets and filaments, could suppress mass accretion onto halos within those sheets and filaments. This mechanism would clearly be incapable of explaining assembly bias in high-mass halos, which collapse before the formation of sheets and filaments (Bond et al. 1996), and Sandvik et al. (2007) found that this mechanism fails in low-mass halos as well. Desjacques (2008) investigated the effects of environment on the Bond & Myers (1996) ellipsoidal collapse model, but could not reproduce the sign reversal of assembly bias between high and low masses, or the magnitude of assembly bias seen in N-body simulations.

In this paper, we will argue that many aspects of assembly bias may be understood quite simply. In particular, we show that the statistics of Gaussian random fields require significant assembly bias in high-mass halos, and we show using N-body simulations that the clustering of massive halos matches precisely the assembly bias that we predict. At low halo masses, we argue that assembly bias arises largely due to the cessation of mass accretion onto a...
HALO ASSEMBLY BIAS

2. BIAS

In this section, we briefly review previous results on biasing and assembly bias. We start by assuming that massive dark matter halos arise from the collapse of peaks in the initial (linear) density field; this allows us to infer many of the properties of halo statistics using the statistics of Gaussian fields (Press & Schechter 1974; Bardeen et al. 1986; Bond et al. 1991; Bond & Myers 1996). High peaks of the density field will in general cluster more strongly than the average matter clustering, which leads to halo bias. In the linear bias model, we directly relate the halo fluctuations to the local matter density fluctuations,

$$\delta_h = b \delta,$$

where $$\delta \equiv \delta \rho / \bar{\rho}$$ is the overdensity and $$b$$ is the bias parameter. On scales over which this expression is valid, we may express halo correlations in terms of the matter correlations, scaled by appropriate powers of the bias. For example, the halo power spectrum would become $$P_h(k) = b^2 P_m(k)$$.

The clustering of halos reflects the clustering of initial peaks only in part, however. Another contribution to halo clustering comes from the motion of peaks along large-scale matter flows. At late times, the halo overdensity becomes, at linear order,

$$\delta_h(a) = \delta_L + \delta_m(a),$$

where the Lagrangian overdensity $$\delta_L$$ corresponds to the clustering of the peaks in the initial density field, and the second term $$\delta_m$$ reflects the motion of the peaks due to advection along matter flows. On sufficiently large scales, in which we can treat the motion of the peaks as tracing the average bulk motion, the second term is simply the matter overdensity. Then the (Eulerian) bias becomes

$$b(a) = \delta_h(a) / \delta_m(a) = 1 + \delta_L / \delta_m(a) = 1 + b_L(a).$$

Note that we explicitly include the time dependence of the matter overdensity, but that the Lagrangian overdensity of peaks is not time dependent: it simply reflects the clustering of peaks in the initial conditions. Therefore, over time, the clustering associated with motion will dominate over the clustering from the initial conditions. That is, if we can treat the halos as test particles, we expect their Lagrangian bias to decay with time inversely with the linear growth factor, $$b_L(a) \propto 1/D(a)$$ (e.g., Fry 1996; Tegmark & Peebles 1998). This behavior holds only in the test particle limit: if the halos change over time, e.g., by growing in mass or by destruction due to incorporation into larger structures, then the bias will not generally evolve as $$b_L(a) \propto 1/D(a)$$ (e.g., Conroy et al. 2007).

Henceforth, we will focus on the Lagrangian bias $$b_L$$, noting that the conversion to Eulerian bias $$b$$ is simply given by equation (3) on linear scales. The next step is to determine $$b_L$$. From equation (1), we can think of bias as the rate at which the number of halos or peaks changes as the background density is varied.

3. SIMULATION AND POSTPROCESSING

We have performed $$N$$-body simulations of structure formation in scale-free cosmologies, with $$\Omega_m = 1$$ and power-law power spectra. We have chosen to run scale-free models rather than more realistic Lambda CDM models in order to simplify the problem as much as possible and to remove possible effects from late-time cosmic acceleration or a varying spectral index. On scales far from the box size and far from the Nyquist frequency, the only relevant scale in the problem is $$M_i(a)$$, the characteristic nonlinear mass scale satisfying $$\delta(M_i, a) = \delta_c = 1.68$$. If the power spectrum slope is $$n$$, such that $$P(k) \propto k^n$$, then the rms mass fluctuations scale like $$\sigma(M) \propto M^{-(3+n)/6}$$. Because the linear growth factor behaves as $$D(a) = a$$ for this critical-density cosmology, the nonlinear mass scales like $$M_i(a) \propto a^{6/(3+n)}$$. Since the nonlinear mass varies rapidly with expansion factor $$a$$, by examining halo properties at multiple redshifts we can study halo properties over a wide range of $$M/M_i$$ (Efstathiou et al. 1988; Wechsler et al. 2006).

In this paper, we will focus on the largest simulation we ran, with 1024$$^3$$ particles, using a power spectral index $$n = -2$$. We normalized the power spectrum such that $$\sigma(M, a) = 10 a M^{1/6}$$, where the mass $$M$$ is measured in number of particles. The simulation was evolved using the adaptive P3M code GRACOS (Shirokov & Bertschinger 2005), using a Plummer softening length $$\varepsilon_p = 0.1$$ times the mean interparticle spacing. The simulation was started at $$a = 0.005$$, and 100 outputs were generated.
spaced evenly in $\log a$ from $a = 0.0825{-}1$, containing phase-space coordinates (comoving positions and peculiar velocities) and IDs for each particle.

The phase space data for each output are used to find the (sub)halos in a two-step process. First we generate a catalog of halos using the friends-of-friends (FoF) algorithm (Davis et al. 1985) with a linking length of $b = 0.168$ times the mean interparticle spacing. This procedure partitions the particles into equivalence classes by linking together all particles separated by a distance less than $b$, with a density of roughly $\rho > 3/(2\pi b^3) \simeq 100$ times the background density. We keep all halos with more than 32 particles.

We identify subhalos of the FoF groups using an independent implementation of the SUBFIND method of Springel et al. (2001) which defines subhalos as locally overdense, gravitationally self-bound aggregations of particles bounded by a density saddle point. After some experimentation with different techniques we found this method gave a good match to what would be selected “by eye” as subhalos. SUBFIND works by first sorting particles by local density, and then building the subhalo population by adding particles in order of decreasing density threshold. We use a spline kernel with 16 neighbors to estimate the density. The particles attached to each peak are considered as a subhalo candidate. When the two nearest neighbors of a particle have larger density than it does, and each belongs to a different subhalo, the union of the two subhalos is also considered as a subhalo candidate. The candidates are then checked for self-boundedness in the inverse order in which they were generated. We call the most massive subhalo in any FoF group the central halo. The other subhalos we refer to as satellites, and we keep all subhalos with more than 20 particles.

For each subhalo we compute and store a number of properties including the bound mass, velocity dispersion, peak circular velocity ($v_c$), total potential energy, position, and velocity. Since mass loss can be quite extreme for some subhalos, we will use $v_c$ rather than mass to quote subhalo “size.” The mass, peak circular velocity, and potential are highly correlated; however, since our subhalos are not spherical, there is a nontrivial scatter between $v_c$ and mass at fixed time.

A merger tree is computed from the set of subhalo catalogs by identifying for each subhalo a “child” at a later time. We process four consecutive simulation outputs at a time and for each subhalo in the earliest output we define its child as the subhalo at a later time which maximizes

$$\alpha = \ln^{-1}\left(\frac{a_2}{a_1}\right)\left(1 - \frac{|M_1 - M_2|}{M_1 + M_2}\right) \sum_{i \leq 1} \phi_i^2,$$

where $a_1$ and $M_1$ are the scale factor and mass of the progenitor, $a_2$ and $M_2$ are the candidate halo at the later time, $\phi_i$ is the potential of particle $i$ computed using the particles in subhalo 1, and the sum is over all particles in the progenitor that also lie in the candidate. The weighting by $\phi_i^2$ ensures that the “core” particles in the progenitor are given more weight than the less-bound particles. We found that weighting by $\phi_i^2$ gave better results than weighting by $\phi_i$, but higher powers of $\phi$ did not perform appreciably better than $\phi_i^2$. We find ~95% of the subhalos have a child in the next time step, with 1% or 2% skipping one or more output times. A similar method for subhalo tracking has been employed by Springel et al. (2005), Faltenbacher et al. (2005), Kravtsov et al. (2004), Allgood et al. (2006), and Harker et al. (2006).

In addition to the Eulerian halo properties mentioned above, we also require Lagrangian properties of halos, such as their centroids or mean initial velocities. From the initial conditions for the simulation, we have the initial density field $\delta(x)$, the initial comoving displacement field $d(x)$ satisfying $\delta + \nabla \cdot d = 0$, the peculiar velocity field, which in the Zel’dovich approximation for $\Omega_m = 1$ is $v = a Hd$, and the strain field $S = -\nabla^2 d$. Note that $\nabla S = \delta$. These fields are all computed on the same grid as the initial (undisplaced) particle positions, by Fourier transform of the initial density field. For example, the strain field $\delta_{S}$ is computed by FFT of $\delta_{kk}/k^2$. We would like to average these quantities over the Lagrangian volumes occupied by the halos. We may do so using the unique IDs for each particle. For each field $F$, we compute the average $\langle F \rangle$ over every halo by a simple average over the halo’s particles:

$$\langle F \rangle_i = N_i^{-1} \sum_{j=1}^{N_i} F(x_{ij}),$$

where $x_{ij}$ is the Lagrangian position of the $j$th particle of halo $i$. Lastly, we will also require derivatives of these smoothed properties with respect to mass, e.g., $d(\delta)/d(\log M)$. We use the merger tree described above to compute derivatives with respect to mass. For every halo, we move along the most massive branch of its merger tree for a factor of 2 in mass, tabulating mass $M$ and the average $\langle F \rangle$. We then fit a linear expansion, $F = F_0 + F_1 \Delta(\log M)$, in order to measure the derivative $F' = d\langle F \rangle/d\log M$.

Lastly, we will determine bias parameters in Fourier space. We measure the matter power spectrum $P(k)$, the halo auto-power spectrum $P_h(k)$, and the halo-matter cross spectrum $P_{hh}(k)$, and define the halo bias as $b = P_h(k)/P(k)$. We use the ratio of the cross-spectrum rather than the halo auto-spectrum, because the former is less sensitive to shot noise in the halo counts.

4. HIGH-MASS HALOS: $M \gg M_*$

We first focus on halos well above the nonlinear mass scale, $M > M_*$, where $\sigma(M) \ll \delta_c$. These halos are particularly simple: because they are so massive, they dominate their environments, and because they are far out on the tail of the probability distribution of $\delta$, the mass dependence of their bias scales like $b \approx \delta_c/\sigma^2$. Antibiasing from larger scale halos should be negligible. Another key simplification in this regime is that the collapse of these rare peaks is nearly spherical (Bardeen et al. 1986, Appendix C). Because the collapse is roughly spherical, we expect the collapse threshold predicted by the spherical collapse model (Gunn & Gott 1972) to be a good approximation: $\delta_c \approx 1.68$. This expectation is borne out by our $N$-body results. In Figure 1, we plot the average linearly evolved overdensity for FoF halos well above $M_*$ and indeed we find $\langle \delta \rangle \approx 1.7$ for the $\gtrsim 3 \sigma$ peaks. Note that multiple different redshifts and halo masses are plotted in this figure.

The validity of $\delta_c \approx 1.68$ for halos in the regime $\sigma \ll 1$ allows us easily to estimate the halo bias as $b \approx \delta_c/\sigma^2$ as mentioned above. This is the average over all peaks of a given mass, which marginalizes over any other peak parameters which could affect bias. One obvious additional parameter which the bias must depend on is the curvature of the peak, which we will parameterize by $S = d(\delta)/d(\log M)$. To see this, note that we can estimate the large-scale background density $\delta_b$ in which the halos reside by a Taylor expansion: $\delta_b \approx \delta + \Delta(\log M) d\delta/d(\log M) + \ldots$. Two peaks of the same mass $M$ and peak height $\delta$ but different peak curvatures $\kappa_1 > \kappa_2$ (i.e., $|\kappa_1| < |\kappa_2|$) will tend to live in different environments: peak 1 will tend to have a larger background density than peak 2, and therefore will have a larger bias. It is straightforward to estimate the dependence of bias on curvature. Writing
\[ \langle \delta \rangle = \delta / \sigma, x = s / \sigma, \text{ and the cross-correlation coefficient } \gamma = \langle \nu x \rangle, \]

the bias becomes (Bardeen et al. 1986)\(^4\)

\[ b_L \approx \sigma^{-1} \frac{\nu - \gamma x}{1 - \gamma^2}. \tag{8} \]

We can easily sketch a derivation of this expression following the argument of Kaiser (1984). Let us compute the correlation function for regions above a high threshold \( \nu \). Consider regions \( 1 \) and \( 2 \), each of size \( R \) and separated by a distance \( r_{12} \), with peak heights \( \nu_1 \) and \( \nu_2 \) and curvatures \( x_1 \) and \( x_2 \). These Gaussian fields are described by a covariance matrix with diagonal elements \( \langle \nu^2 \rangle = \langle x^2 \rangle = 1 \). Of the off-diagonal elements, only \( \langle \nu_1 \nu_2 \rangle = \psi(r_{12}) \) and \( \langle \nu_1 x_1 \rangle = \langle \nu_2 x_2 \rangle = \gamma \) are important; long-range correlations involving derivatives of the density field fall off more quickly than these terms by powers of \( R/r_{12} \). We compute the two-point correlation of regions with \( \nu > \nu_1 \) by integrating the Gaussian probability over \( \nu_1 > \nu \) and \( \nu_2 > \nu \). If we are uninterested in the curvature \( x \), then we first marginalize over \( x_1 \) and \( x_2 \), leaving a \( 2 \times 2 \) covariance matrix with off-diagonal element \( \psi \), and we find that the two-point function for regions above threshold becomes \( \xi_{pk} = \nu^2 \psi \), giving a Lagrangian bias \( b_L = \nu \sigma / \psi \) (Kaiser 1984). We can similarly compute the two-point function for regions above threshold with a specified curvature \( x \) by repeating this calculation, but not marginalizing over \( x_1 \) and \( x_2 \). Recalling that the one-point probability distribution for \( \nu \) at a given \( x \) is a Gaussian centered at \( \langle \nu | x \rangle = \gamma x \) with variance \( 1 - \gamma^2 \), we change variables from \( \nu \) to \( \nu' = (\nu - \gamma x)(1 - \gamma^2)^{1/2} \). Noting that \( \langle \nu_1 \nu_2 \rangle \approx \psi(1 - \gamma^2) \), we immediately see that the two-point correlation function for regions above threshold becomes \( \xi_{pk} = \nu^2 \psi(1 - \gamma^2) \), giving a bias \( b_L = \nu / \sigma(1 - \gamma^2)^{1/2} \), i.e., equation (8).

The bias of halos found in our simulation matches equation (8) quite well. For a power spectrum index \( n = -2 \), as we have used, note that \( \sigma_x = \sigma / \sqrt{6} \) and \( \gamma = -1/\sqrt{6} \), for a top-hat window function. In Figure 2 we plot the halo bias as a function of peak curvature, along with our prediction from equation (8). There is evidently a strong dependence of bias on peak curvature. The \( \sim 40\% \) range in \( \delta \) corresponds to a factor of 2 variation in the correlation function for halos of this mass. One subtlety which we have glossed over in computing the statistics of high-mass peaks is the issue of the appropriate smoothing to use. We have assumed a top-hat smoothing; however, there is no compelling theoretical reason to favor a top-hat kernel versus other possibilities, such as a Gaussian (although the sharp \( k \)-space filter implicit in the Press-Schechter model appears unphysical). As a sanity check on our top-hat smoothing, we have computed the average Lagrangian volumes corresponding to high peaks, by stacking roughly 1000 halos of mass \( \sim 3000M_\odot \). The results are shown in Figure 3. The average Lagrangian volume occupied by FoF halos is reasonably compact, with the probability for halo inclusion steeply varying from near 1 to near 0 over a range \( 0.8 < r_{\text{TH}} < 1.2 \), where \( r_{\text{TH}} \) is the top-hat smoothing radius for mass \( M \). In comparison, a Gaussian smoothing window is clearly more diffuse. We conclude that our use of top-hat smoothing is reasonable, although we caution that these results may depend on the slope of the power spectrum. Finally, we note that there is a nontrivial amount of mass inside \( r < r_{\text{TH}} \) that does not end up in the halo; we will return to this point below in § 5.

\(^4\) Bardeen et al. (1986) parameterized curvature not with \( d\delta / d \log M \) as we do, but rather with the closely related \( \nabla^2 \delta \).
We have shown that there is a strong dependence of halo clustering on peak curvature. How does this relate to the assembly bias observed in previous simulations?

In the high-mass regime, there is a direct relationship between peak curvature and halo assembly history. Recall that in the spherical collapse model, regions complete their collapse once the average linearly evolved overdensity \( \delta \) reaches \( \delta_c \). This implies that we can estimate the assembly history of halos simply by measuring how the smoothed overdensity \( \delta \) varies with smoothing scale. We would expect the accretion rate and the curvature to be related by

\[
\frac{d (\log M)}{d (\log a)} \approx \frac{C_0}{d (\log \delta_c / d (\log M))}.
\]

The average mass accretion histories for high-mass halos appears consistent with this expectation. In Figure 4 we plot the stacked mass accretion history profiles for halos of 120 particles at \( a = 0.1089 \), splitting the population based on peak curvature. The average accretion rates are in good agreement with our expectations. This is true only in the average, however: individual accretion histories are noisy because much of the accretion onto halos is not smooth, but rather consists of discrete accretion events. This may explain why different age indicators for halos give different levels of assembly bias (Wetzel et al. 2007). The noisiness of commonly used age indicators may be large enough to mask the strong assembly bias clearly present in high-mass halos.

Another halo parameter seen to correlate with bias is concentration (Wechsler et al. 2006). While we cannot write down an expression directly relating peak curvature and halo concentration, we nevertheless would expect curvature and concentration to be correlated. Previous work has shown that concentration and peak curvature may be correlated. Physically, we would expect that highly curved peaks should produce more concentrated halos. To test this, we plot in Figure 5 the stacked radial profiles of halos with 4000–5000 particles at \( a = 0.2501 \) as a function of Lagrangian peak curvature. There is a correlation between initial peak curvature and final halo concentration apparent in the stacked profiles, although it is quite apparent that peak curvature is not
equivalent to concentration. While the inner profiles ($r < r_s$) of curvature-selected halos match well with profiles of concentration selected halos, the outer profiles ($r > r_s$) begin to diverge.

Within the statistical errors of our finite sample of rare halos, we have found no residual dependence of bias on concentration, once we account for the peak curvature dependence described above. Our simulation had relatively low force resolution, and so our fitted concentrations may be sufficiently noisy to mask any underlying residual concentration dependence of bias, but it appears that assembly bias in rare, high-mass halos is in agreement with theoretical expectations. The interpretation of this effect is straightforward: halos residing in high-density environments have an enhanced mass supply, and therefore grow faster, than halos in low-density environments.

5. LOW-MASS HALOS: $M \ll M_*$

In this section we consider assembly bias for low-mass halos with $M \ll M_*$, for which $\sigma(M) \gg \delta_i$. The age and concentration dependence of bias at low masses is considerably stronger than in the high-mass regime, and of the opposite sign: the oldest, highly concentrated halos have a bias more than twice as large as that of the youngest, low-concentration halos of similar mass $M \ll M_*$ (Gao et al. 2005; Wechsler et al. 2006; Gao & White 2007).

In agreement with previous workers, we also find this reversed sense of assembly bias. In keeping with our Lagrangian viewpoint, we use mean Lagrangian overdensity $\langle \delta \rangle$ as a proxy for age. In Figure 6, we plot average accretion histories for halos of 200 particles at $a = 1$, split into quartiles of $\langle \delta \rangle$. The oldest halos show little growth at late times, accreting only $\sim 20\%$ of their mass after $a = 0.5$. In comparison, the nonlinear mass scale increases by a factor $\sim 64$ over this same time.

Because the oldest halos do not grow appreciably in mass, we can think of their dynamics as those of test particles following large-scale bulk flows. As discussed in § 2, we therefore expect the bias of this population to relax over time to $b \rightarrow 1$. This is indeed the behavior observed for our oldest halos, as shown in Figure 7, which plots the (Eulerian) bias of the same halos shown in Figure 6. The oldest $\sim 20\%$ of halos (i.e., the rightmost bin) are nearly unbiased, while the remainder are strongly antibiased, to a degree consistent with predictions of even the simplest Press-Schechter model, $b \sim 1 - \delta_i^{-1} \approx 0.4$ (Cole & Kaiser 1989).

The magnitude of assembly bias present at low masses is therefore not surprising. Most low-mass halos are antibiased at expected levels, while a subpopulation becomes unbiased. The theoretical challenge is to explain the number of unbiased, low-mass halos.

One hint toward the origin of the nonaccreting, unbiased population comes from the observation that many, if not most of these halos are associated with nearby larger halos (e.g., Wang et al. 2007). In Figure 7, we again plot the bias of $\sim 200$ particle halos, but excluding those which in earlier outputs were subhalos. This removes $\sim 20\% - 25\%$ of the total population. Because the time sampling of our simulation outputs do not finely cover the dynamical times of all halos, this cut could miss subhalos with orbital pericenters inside of larger halos, but apocenters well outside. Ludlow et al. (2008) have recently argued that a significant fraction of subhalos have orbital apocenters extending out to many times the virial radius of their host halo. To account for this, we also conservatively exclude all halos which are within $3r_{vir}$ and gravitationally bound to larger halos. This will exclude the population discussed by Ludlow et al. (2008) as well as any halos entering the infall region of massive objects. This cut removes a further $\sim 15\%$ of low-mass halos, leaving $\geq 60\%$ of the original low-mass population. The effect of these cuts is effectively to exclude the old subpopulation, even though the cuts do not explicitly refer to halo age or accretion history. As can be seen in Figure 7, the old (high $\delta$) bin is essentially emptied, and the bias of the remaining halos is similar to that obtained before applying this preselection (compare red and blue curves). It therefore does seem likely that, as suggested by previous authors, the old subpopulation of low-mass halos corresponds to early formers whose
growth was stunted by environmental effects (Diemand et al. 2007; Wang et al. 2007).

One plausible environmental effect that would account for the arrested development of the oldest low-mass halos is simply the enhanced velocity dispersion in the vicinity of massive halos. In environments where the ambient velocity dispersion greatly exceeds the escape velocity of low-mass halos, those halos will be unable to accrete significant material. We indeed find that the oldest halos preferentially reside within hot environments, as shown in Figure 8.

Given that the oldest low-mass halos tend to exist in proximity to larger halos, we would expect them to originate (in Lagrangian space) from the vicinities of high-mass halos as well, cf. Figure 3. As discussed in § 4.1, much of the material within the Lagrangian volumes of high-mass halos ends up within the final collapsed halo, so how do we understand the survival of the low-mass halos in the presence of their rapidly accreting neighbors? One possible explanation noted by Ludlow et al. (2008) is that these halos are ejected out of larger halos by three-body interactions. Because our goal is to understand the clustering of halos in terms of their Lagrangian properties of known (Gaussian) statistics, and since complicated multibody interactions are clearly difficult to model using only Lagrangian quantities, we have not attempted to address this mechanism using our simulations. Rather, we have tried to find Lagrangian properties of halos which correlate with their survival around larger structures.

One Lagrangian quantity which correlates with survival is plotted in Figure 9. We have taken halos of mass $40,000–60,000$ particles at $a = 1$, nearly $M \simeq M_\star$, and identified smaller scale peaks within the spherical Lagrangian volumes centered on the larger halos. As shown in the figure, the relative binding energy of the subpeaks roughly correlates with their escape likelihood.

We have defined the binding energy $\delta E$ in the following manner. The gravitational potential near a peak may be written

$$\phi(x) \approx \phi_0 + x \cdot \nabla \phi + \frac{1}{2} x \cdot \nabla \nabla \phi \cdot x + \ldots$$

and the total perturbed energy becomes $\delta E = \delta \phi + \delta T$. In computing $\delta \phi$, we smooth the strain tensor $S = \nabla \nabla (\nabla^{-2} \delta)$ to the potential’s curvature tensor. We can similarly write the kinetic energy near the peak as

$$T(x) = \frac{1}{2} |v|^2 = \frac{1}{2} \left[ aH(x + d) + v_p \right]^2
\approx \frac{1}{2} \left( aH |x|^2 + 2(aH)^2 x \cdot d + \ldots \right),$$

and the unperturbed piece again does not affect the peak dynamics, and so we write the perturbed kinetic energy as

$$\delta T = 2(aH)^2 x \cdot d,$$

where we have used the Zeldovich approximation for $\Omega_m = 1$ to relate the peculiar velocity and comoving displacement, $v_p = aHd$. The unperturbed piece again does not affect the peak dynamics, and so we write the perturbed kinetic energy as

$$E' \equiv \delta E + (aH)^2 (2r^2 - r_{wm}^2) = 0$$

roughly separates accreted subpeaks from escaping subpeaks.

Our assumption, therefore, is that subpeaks inside of larger peaks can escape accretion onto their hosts if the relative energy
is large enough, $E' > 0$. Due to environmental effects, however, they are unable to grow, and thereby produce the population of old, unbiased low-mass halos. In §6, we test how well this simple picture can reproduce our numerical bias results.

6. A TOY MODEL FOR ASSEMBLY BIAS

In §§4 and 5, we have discussed assembly bias in two asymptotic regimes, $\sigma \ll 1$ and $\sigma \gg \delta_c$. We found that two distinct effects generated assembly bias in these two regimes. At intermediate masses, $\sigma \sim 1$, both effects can be significant, and the overall bias of a sample of halos will involve a competition between these two effects. The transition between the high-mass behavior and the low-mass behavior will depend in detail on the precise selection criteria used to define the sample of halos. For example, we plot in Figure 10 the halo bias for halos selected on peak height $h_i / \delta_c$ and on peak curvature $d \log (\delta) / d \log M$. Recall that these peak properties both affect halo properties like age and concentration. At high masses, the scatter in peak height among peaks above threshold is small, and so the scatter in peak curvature dominates the assembly bias. At low mass, the peak height is the dominant factor determining halo age or concentration, causing a reversal in the sense of assembly bias. To illustrate, we plot in the right panel of Figure 10 the bias for halos selected on $\delta / \delta_c - d \log (\delta) / d \log M$. At high masses, the curvature term dominates, while at low masses, the first term dominates. Halo samples selected on formation time or concentration show similar behavior: like the right panel of Figure 10, at high masses the curvature dependence of bias dominates, whereas at low masses, the scatter in peak height dominates the scatter in bias, leading to a reversal in sign for the formation time or concentration dependence of bias (Wechsler et al. 2006; Gao & White 2007). At intermediate masses, both of these effects are significant and

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**Fig. 9.**—Binding energy vs. initial radius for peaks that are accreted (red points), that become subhalos (green points), or that escape (blue points).

**Fig. 10.**—Bias as a function of peak parameters. Bias is plotted for halos in the range 200–300 particles at various redshifts. **Left:** Red and blue curves show bias for the upper and lower 20% of peak height $h_i / \delta_c$, where $\delta_c$ is the ellipsoidal collapse threshold calculated from each halo’s local strain tensor. **Middle:** We split halos on peak curvature, $d \log (\delta) / d \log M$. **Right:** We split halos on an average of these two quantities, $(\delta / \delta_c - d \log (\delta) / d \log M)$. At high masses, the scatter in curvature dominates, while at low masses, the scatter in peak height dominates, leading to a transition in assembly bias at intermediate masses. The dependence of bias on formation time or concentration shows similar behavior as the right panel, for essentially the same reason (see text).
compete with each other, and the magnitude and sign of assembly bias will depend in detail on the precise age indicator that is employed. For example, an age indicator that correlates more strongly with peak height rather than peak curvature will show a crossover at higher masses than the example shown in Figure 10c. This may explain why different measures of halo age give different amounts of assembly bias when applied to the same samples of halos (Li et al. 2008).

Using our results on the assembly bias of low- and high-mass halos, we have constructed a simple toy model for halo biasing. We generate a Gaussian random density field, smooth the field on multiple scales, and starting from the largest smoothing scales, search for density peaks exceeding the ellipsoidal collapse threshold (Bond & Myers 1996, hereafter BM96). Isolated peaks above threshold are labeled as collapsed halos, whereas peaks falling within 1.2 smoothing radii (e.g., Fig. 9) of larger scale collapsed regions are only labeled as collapsed if their binding energies relative to their larger hosts exceeds the threshold found empirically from our N-body halos, i.e., equation (13).

A key component of this calculation is the collapse threshold. A comparison of the simple BM96 model with our N-body results shows that the model works well at predicting collapse thresholds for reasonably high peaks, \( \sigma \approx 1 \); however, at low peaks \( \sigma \gg 1 \), the model begins to break down. Very roughly, the ellipsoidal collapse threshold is larger than the spherical collapse threshold by a ratio \( \delta e_c / \delta e_\infty \approx \left( 1 - \frac{1}{2} \delta e_\infty \right)^{-1} \), where the ellipticity \( e_i \) of the strain tensor is given by the difference between its largest and smallest eigenvalues, \( e_i = \left( S_{33} - S_{11} \right) / 2 \delta \). The ellipticity and prolaticity are random variables whose probability distribution is related to peak height (Bardeen et al. 1986; BM96), but very roughly the mean ellipticity scales as \( \langle e_i \rangle \sim 1/(2 \nu) \). For \( \nu < 1 \), the BM96 collapse threshold quickly grows, and for sufficiently large \( e_i \) and \(-p_s\), the model predicts that collapse never occurs. However we find many low-mass halos with average peak heights \( \langle \delta \rangle \) falling far below the collapse thresholds expected given their average Lagrangian strain tensors \( \langle S \rangle \). This is not surprising. As noted explicitly by BM96, their homogeneous ellipsoid model assumes that the evolution of the external tidal field acting on the collapsing peak may be determined from the strain tensor \( \langle S \rangle \) averaged over the local volume of the peak. While it is certainly true that the external tidal field and local strain tensor are closely related at early times, in the nonlinear regime the one-to-one correspondence between them breaks down. On scales where \( \sigma \gg 1 \), the external tidal field evolves nonlinearly in a manner that cannot be predicted from purely local measurements over the peak volume. Therefore, we would not expect the simple BM96 model to apply in this regime, as those authors note.

Nevertheless, we require some choice of collapse threshold for our toy model. At low masses, the lowest \( \langle \delta \rangle \)'s are roughly \( \approx 2.5 \) (see Fig. 7), so we arbitrarily set a threshold \( \delta_c = \min(\delta_{\text{BM96}}, 2.5) \). In addition, we also require that subpeaks collapsing inside of larger peaks must complete their collapse sufficiently before the larger peak collapses. We can estimate the collapse redshift \( z_c \) by \( 1 + z_c = (\delta_c / \delta_c) \). Writing \( z_c \) and \( z_p \) as the collapse redshifts for the subpeak and large peak, respectively, we require \( (1 + z_c) \geq 2(1 + z_p) \) for a subpeak to be labeled as a collapsed halo.

Figure 11 shows results of this simple calculation. In order to compare with Figure 10, we plot bias as a function of the same peak parameters. In this figure, the Lagrangian bias \( b_L \) of the peaks was computed, and the Eulerian bias was assumed to behave as \( b = 1 + b_t \). At high masses, \( \sigma < 1 \), the bias matches equation (8), as it must. At low masses (\( \sigma \gg 1 \)) the bias of isolated peaks asymptotes to \( b_E \approx 0.4 \), while the subpeaks become unbiased, \( b_E \rightarrow 1 \). These features appear fairly generic. Behavior in the transition region, \( \sigma \sim 1 \), depends on details of the model. In particular, the bias of the old halos in the intermediate mass regime depends on the fraction of halos that are subpeaks, and varying parameters in the model changes this fraction.

Overall, however, we find reasonable agreement between the toy model and our N-body simulation in the regimes where assembly bias is significant, for \( M \gg M_* \) and \( M \ll M_* \). The point of this exercise is not to claim that we have solved the halo formation problem, but merely to illustrate that the assembly bias in these regimes is generic. The assembly bias at high masses is an unavoidable consequence of Gaussian statistics, and therefore any halo formation model which relates initial density peaks to halos will also find similar levels of assembly bias, depending somewhat on the peak definition. The behavior at low masses may appear more model dependent; however, we do not believe our results are specific to our particular mechanism for production of nonaccreting halos. Rather, we expect that any mechanism which produces roughly the correct fraction of nongrowing, low-mass halos will also match the assembly bias found in N-body simulations.

7. DISCUSSION AND CONCLUSIONS

We have investigated halo assembly bias in hierarchical structure formation. We find that assembly bias is significant both at high masses (\( M \gg M_* \)) and low masses (\( M \ll M_* \)). Two different competing effects generate assembly bias in these two regimes. At high masses, peaks of low curvature are more strongly clustered than peaks of large curvature, in a manner entirely consistent with the statistics of peaks of random Gaussian fields. We provide a simple expression for the bias in the high-mass regime which accounts for assembly bias and reduces to the Press-Schechter bias when marginalized over all parameters excluding mass (note that fitting formulae for bias calibrated at lower mass, such as those proposed by Sheth & Tormen [1999] or Seljak & Warren...
[2004] fail in the high-mass regime (Cohn & White 2007]). At low masses, assembly bias appears largely driven by the formation of a nonaccreting subpopulation of low-mass halos in the vicinity of larger halos. Because the nonaccreting halos become unbiased, they are more strongly clustered than accreting halos of similar mass, which are antibiased. We found that a simple toy model is able to reproduce our numerical bias results in the asymptotically large and small mass regimes. At intermediate masses ($M \approx M_\text{c}$), both of the above effects become important, and the assembly bias for a sample of halos will involve a weighted average of both effects.

The application of these results to observed objects like galaxies or quasars is not straightforward, because it will depend on the halo occupation distribution (HOD) of those objects. The very simplest HOD models assume that the galaxy content depends only on halo mass, and not on assembly history, whereas semianalytic models used in conjunction with N-body simulations tend to find a strong dependence of galaxy properties on assembly history at fixed mass (e.g., Springel et al. 2005; Croton et al. 2007). If the former assumption is valid, then assembly bias will not be important for galaxy clustering. Perhaps surprisingly, Tinker et al. (2008) have recently found that properties of observed SDSS galaxies do not appear to correlate significantly with local environment, once the dependence on host halo mass has been accounted for, suggesting that assembly bias may indeed be unimportant for modeling galaxy clustering (see also Blanton et al. 2006; Yang et al. 2006; Berlind et al. 2006; Blanton & Berlind 2007; Wang et al. 2008; Skibba & Sheth 2008).

Besides the clustering of halos, our results also have implications for our understanding of halo assembly. In particular, we have found that massive halos are very well described by spherical collapse. The average precollapse Lagrangian volumes occupied by the most massive halos are approximately spherical regions of size $r \sim r_{\text{TH}} = (3M/4\pi\rho_0)^{1/3}$. However, $r_{\text{TH}}$ is not a sharp boundary in Lagrangian space: there is significant mass outside $r > r_{\text{TH}}$ at early times which is found inside the virial radius at late times, and similarly there is considerable mass at $r < r_{\text{TH}}$ which avoids accretion at late times. Indeed, we might speculate that the postcollapse radial distribution of the matter inside and around a dark matter halo is simply related to the precollapse profile of local density and potential. Since both of these quantities are described by well-understood Gaussian statistics, this raises the intriguing possibility of understanding the internal structure of virialized objects from first principles. This topic is the subject of further study.

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