First order strong convergence and extinction of positivity preserving logarithmic truncated Euler-Maruyama method for the stochastic SIS epidemic model

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Abstract
The well-known stochastic SIS model characterized by highly nonlinear in epidemiology has a unique positive solution taking values in a bounded domain with a series of dynamical behaviors. However, the approximation methods to maintain the positivity and long-time behaviors for the stochastic SIS model, while very important, are also lacking. In this paper, based on a logarithmic transformation, we propose a novel explicit numerical method for a stochastic SIS epidemic model whose coefficients violate the global monotonicity condition, which can preserve the positivity of the original stochastic SIS model. And we show the strong convergence of the numerical method and derive that the rate of convergence is of order one. Moreover, the extinction of the exact solution of stochastic SIS model is reproduced. Some numerical experiments are given to illustrate the theoretical results and testify the efficiency of our algorithm.

Keywords: Stochastic SIS epidemic model; Logarithmic truncated Euler-Maruyama method; Positivity-preserving; First order convergence rate; Extinction

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1. Introduction
Infectious diseases are the public enemy of human population and seriously threatening the health and life. In many research methods of infectious disease, epidemic models are widely used for analyzing the spread and control of infectious diseases. A common type of mathematical models in epidemiology is known as the SIS (Susceptible-Infected-Susceptible) epidemic models, which are used to describe some infections that don’t have permanent immunity and individuals become susceptible again after infection. For example, Hethcote-Yorke [1] has indicated that rota-viruses, some sexually transmitted and bacterial infections have such characteristics. In the real world, epidemic spreads are always affected by demographic and environmental stochasticity. However, deterministic model can’t capture these fluctuating features. Hence, the stochastic SIS epidemic model

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[2], which can provide description and prediction more accurately than that of their deterministic counterpart, have attracted considerable attention and is described by

\[
\begin{align*}
\frac{dS(t)}{dt} &= \left[\mu N - \beta S(t)I(t) + \gamma I(t) - \mu S(t)\right]dt - \sigma S(t)I(t)dB(t), \\
\frac{dI(t)}{dt} &= \left[\beta S(t)I(t) - (\mu + \gamma)I(t)\right]dt + \sigma S(t)I(t)dB(t),
\end{align*}
\]

with initial values \(I(0) + S(0) = S_0 + I_0 = N\), where \(S_t\) denotes the number of susceptibles, \(I_t\) is the number of infecteds at time \(t\) and \(N\) is the total amount of entire population among whom the disease is spreading. Moreover, \(\mu \geq 0\) is the per capita death rate, and \(\gamma \geq 0\) is the cure rate of infected individuals, so \(1/\gamma\) is the average infectious period. In addition, \(\beta \geq 0\) is the disease transmission coefficient, and \(B(t)\) is a standard Brownian motion defined on a complete probability space. \(\sigma\) is the variance of the number of potentially infectious contacts that a single infected individual makes with another individual over the small time interval \([t, t + dt]\).

It is worth noting that it is easy to get \(I(t) + S(t) = N\) by solving

\[
\frac{d[S(t) + I(t)]}{dt} = \left[\mu N - \mu (S(t) + I(t))\right]dt.
\]

Therefore, (1.1) can be written equivalently as follows

\[
\frac{dI(t)}{dt} = \left[\eta I(t) - \beta I(t)^2\right]dt + \sigma I(t)(N - I(t))dB(t),
\]

with initial value \(I(0) = I_0 \in (0, N)\), where \(\eta = \beta N - \mu - \gamma\), and Gray et al. [2] have studied a number of nice properties for the special stochastic SIS epidemic model (1.2). Particularly, here we mention an important theorem from [2], which shows that the unique solution of (1.2) not only exists but also remains within the bounded domain \((0, N)\).

**Theorem 1.1.** For any given initial value \(I_0 \in (0, N)\), the SDE (1.2) has a unique global solution taking values in the domain \((0, N)\), i.e.,

\[
P\{I(t) \in (0, N) \forall t \geq 0\} = 1.
\]

Nevertheless, the true solution of the nonlinear SDE (1.2) is not available, which means that looking for well-performed numerical methods is necessary to simulate and help to observe the behavior of the underlying SIS epidemic model. The most common numerical method for SDEs with global Lipschitz continuous coefficients is classical Euler-Maruyama (EM) method. But Hutzenthaler-Jentzen-Kloeden [3, Theorem 2.1] have showed that, for SDEs whose coefficients are growing superlinearly, EM approximations would inevitably involve the \(p\)th moments diverging to infinity. In recent years, several modified EM schemes, including implicit schemes and explicit schemes, have been proposed for approximating the exact solutions of nonlinear SDEs, see, e.g., [4–14] and the references therein. Even so, these methods mentioned above with requiring a global monotonicity condition cannot be directly used for (1.2).
On the other hand, constructing appropriate numerical approximations, which can preserve the properties of the exact solution as well as reproduce the long-time behaviors for SDEs, has been one of the research focuses in numerical analysis field. For example, Li-Yang [15] have applied a positivity-preserving truncated EM (TEM) method to the stochastic logistic model, and further realized long-time dynamical properties, including the stochastic permanence, extinctive and stability in distribution. Recently, Yi-Hu-Zhao [16] have presented a positivity-preserving logarithmic EM scheme for general SDEs whose solutions are positive. Mao-Wei-Wiriyakraikul [17] have established a new positive preserving TEM method for stochastic Lotka-Volterra competition model. Moreover, there are few results of the positivity-preserving numerical methods for stochastic SIS model (1.2). Specifically, Chen-Gan-Wang [18] have constructed the Lamperti smoothing truncation scheme for (1.2), which can preserve the domain of the original SDEs and have a convergence rate of order 1. And then Yang-Huang [19] have developed a positivity preserving logarithmic EM method, which can be regarded as an improvement on the [18]. Nonetheless, to the best of our knowledge, [15–19] has not yet provided the results of the long-time dynamical properties of the numerical approximation for stochastic SIS model (1.2). Precisely speaking, there is still no result about the numerical methods which not only preserve the positivity but also the long-time dynamical properties for stochastic SIS model (1.2). Thus, we in the paper are devoted to presenting an explicit positivity-preserving numerical scheme and show the extinction of the numerical method.

Inspired by the works in [18, 20, 21], we make use of the Lamperti transformation to construct a type of logarithmic-type transformed numerical method for (1.2). To be more specific, applying the Itô formula to $y(t)$ in the Lamperti-type transformation

$$y(t) = \log(I(t)) - \log(N - I(t)) = \log \frac{I(t)}{N - I(t)},$$

(1.3)

we can arrive at

$$dy(t) = F(y(t))dt + \sigma NdB(t), \quad t \geq 0,$$

(1.4)

with $y(0) = \log(I_0) - \log(N - I_0)$, where $F(x) = \eta - (\mu + \gamma)e^x + \frac{\sigma^2 N^2}{2} - \frac{\sigma^2 N^2}{1 + e^x}$. Note that (1.3) also means

$$I(t) = N - \frac{N}{1 + e^{y(t)}} = \frac{Ne^{y(t)}}{1 + e^{y(t)}},$$

(1.5)

More recently, Li-Mao-Yang [5] proposed a new explicit method and obtained the asymptotic stability of the method under more relaxed conditions. However, the transformed SDE (1.4) is still not satisfied with some conditions in [5]. To deal with transformed SDE (1.4) with exponentially growing coefficients, in this paper, we establish a logarithmic TEM scheme, which can be transformed back to an positivity-preserving approximation of the original SDE (1.2) according to (1.5). This allows us to preserve the same domain $(0, N)$ of the original SDE (1.2) and obtain the first
order rate of convergence. Moreover, one of our important contributions in our paper is the approximation of long-time dynamical properties for (1.2), namely, the extinction of the numerical method is given.

The rest of the paper is structured as follows. Section 2 introduces some preliminary results on certain properties for the stochastic SIS model. Section 3 constructs an explicit scheme preserving the positivity and the domain of the exact solution, and yields the convergence rate of first order. Section 4 focuses on the analysis the extinction of the numerical scheme, i.e. the explicit scheme preserving the extinction of the exact solution of stochastic SIS model. Section 5 provides a couple of numerical examples and simulations to demonstrate the efficiency of the numerical scheme.

2. Mathematical notations and lemmas

Throughout this paper, we use the following notations. Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a complete probability space, and $\mathbb{E}$ denotes the expectation corresponding to $\mathbb{P}$, and $|\cdot|$ denote the Euclidean norm in $\mathbb{R}$. For any $a, b \in \mathbb{R}$, $a \lor b := \max\{a, b\}$, and $a \land b := \min\{a, b\}$. If $\mathbb{D}$ is a set, its indicator function is denoted by $I_{\mathbb{D}}$, namely $I_{\mathbb{D}}(x) = 1$ if $x \in \mathbb{D}$ and 0 otherwise. Suppose $\{\mathcal{F}_t\}_{t \geq 0}$ is a filtration defined on this probability space satisfying the usual conditions (i.e., it is right continuous and $\mathcal{F}_0$ contains all $\mathbb{P}$-null sets) such that $B(t)$ is $\mathcal{F}_t$ adapted. First, we state some useful lemmas which can be found in [18].

Lemma 2.1 ([18, Theorem 3.2]). For any $p \geq 0$ we have
\[
\left( \sup_{t \in [0, T]} \mathbb{E} \left[ I(t)^{-p} \right] \right) \lor \left( \sup_{t \in [0, T]} \mathbb{E} \left[ (N - I(t))^{-p} \right] \right) \leq K_p,
\]
where
\[
K_p = \left( I_0^{-p} \lor (N - I_0)^{-p} \right) \exp \left( pT(|\beta N - \mu - \gamma| + 2\beta N) + \frac{p(p + 1)}{2} \sigma^2 N^2 T \right).
\] (2.1)

Lemma 2.2 ([18, Proposition 3.3]). For any $p \in \mathbb{R}$, the transformed SDE (1.4) has the following exponential integrability property
\[
\sup_{t \in [0, T]} \mathbb{E} \left[ e^{pq(t)} \right] \leq N^{kp} K_p,
\]
where $K_p$ is given by (2.1).

Next we propose EM method to approximate the exact solution of the SDE (1.4). Given a step size $\Delta > 0$ and let $t_k = k\Delta$ for any integer $k \geq 0$. For convenience, we will further use $C_u$ and $C$ to stand for two generic positive real constants dependent on $I_0, \beta, N, \mu, \gamma, \sigma, p$ but independent of the step size $\Delta$ and $k$, which we will use in this section and later sections. Please note that the
values of $C_u$ and $C$ may change between occurrences, where $C_u$ is dependent on $u$. For SDE (1.4) define the EM scheme by
\[
\begin{cases}
Y_0 = y(0), \\
Y_{k+1}^E = Y_k^E + F(Y_k^E)\Delta + \sigma N\Delta B_k,
\end{cases}
\] (2.2)
for any integer $k \geq 0$, where $\Delta B_k = B(t_{k+1}) - B(t_k)$. Transforming back, i.e.
\[
\bar{Y}_{k+1}^E = \frac{N - \frac{N}{1 + e^{Y_k^E}}}{1 + e^{Y_k^E}},
\] (2.3)
gives a strictly positive approximation of the original SDE (1.2).

Lemma 2.3 ([15, Lemma 3.2]). For any $L > 0$ and integer $m \geq 0$, we have
\[
E\left[e^{\frac{|\Delta B_k|^2}{4\Delta}}|\Delta B_k|^{3m}\right] = \sqrt{\frac{2}{\pi}}\Gamma\left(\frac{3m + 1}{2}\right)(4\Delta)^{\frac{3m}{2}}, \quad E[\Theta_k^m F(t_k)] \leq C\Delta^{\frac{3m}{2}},
\] (2.4)
where $\Gamma(\cdot)$ is the Gamma function and $\Theta_k := e^{L|\Delta B_k|}(\Delta^3 + |\Delta B_k|^3)$.

Lemma 2.4 ([19, Theorem 2.1]). For any $p > 0$ there exists a constant $C > 0$ such that
\[
E\left[\sup_{k=0,\ldots,[T/\Delta]} |Y_k^E - y(t_k)|^p\right] \leq C_T \Delta^p
\] (2.5)
for any $\Delta \in (0, 1]$ and $T > 0$, where $[T/\Delta]$ represents the integer part of $T/\Delta$.

Remark 2.5. Although Lemma 2.4 shows that the logarithmic EM method can strongly converge to the exact solution of stochastic SIS model (1.2), which can be found in [19, Theorem 2.2], it has not yet provided the results about the approximation of the long-time dynamical properties for stochastic SIS model (1.2). The main reason is that the diffusion term and nonlinear term of SDE (2.2) will lead to extraordinary values. In other words, a numerical method which can efficiently prevent producing the extraordinary values would be able to preserve the properties of the exact solution on the basis of Taylor expansion. Thus, in the next section, we define the explicit scheme by using an appropriate truncation map.

3. Convergence rate of the logarithmic TEM method

In this section, we aim to construct a logarithmic TEM method, which can preserve the positivity of the original stochastic SIS model (1.2). Moreover, this allows us to preserve the same domain $(0, N)$ of the original stochastic SIS model (1.2) and obtain the first order rate of convergence. For any given stepsize $\Delta \in (0, 1)$, define a TEM scheme by
\[
\begin{cases}
\bar{Y}_{k+1} = Y_k + F(Y_k)\Delta + \sigma N\Delta B_k, \\
Y_{k+1} = \bar{Y}_{k+1} \land \log(K\Delta^{-\frac{1}{2}}),
\end{cases}
\] (3.1)
for any integer \( k \geq 0 \), where \( K > 1 + e^{Y_0} \). Transforming back, i.e.

\[
\tilde{I}_k^\Delta = N - \frac{N}{1 + e^{Y_k}} = \frac{Ne^{\gamma_k}}{1 + e^{\gamma_k}}, \quad \tilde{I}_k = N - \frac{N}{1 + e^{\gamma_k}} = \frac{Ne^{\gamma_k}}{1 + e^{\gamma_k}},
\]

and \( \tilde{I}_k^\Delta \) gives a strictly positive approximation of the original SDS (1.2). Lemmas 3.1 and 3.2 will play their important roles in the proof of the strong convergence of the numerical solutions.

Lemma 3.1. For any \( p > 0 \), the TEM scheme defined by (3.1) has the property that

\[
\sup_{\Delta \in (0,1)} \sup_{0 \leq k \leq [T/\Delta]} \mathbb{E}[e^{pY_k}] \leq \sup_{\Delta \in (0,1)} \sup_{0 \leq k \leq [T/\Delta]} \mathbb{E}[e^{\gamma_k}] \leq CT
\]

for any \( T > 0 \).

Proof. We know that

\[
\bar{Y}_{k+1} \leq Y_k + (\eta + \sigma N^2/2)\Delta + \sigma N \Delta B_k \leq \bar{Y}_k + (\eta + \sigma N^2/2)\Delta + \sigma N \Delta B_k.
\]

Then

\[
\bar{Y}_k \leq Y_0 + (\eta + \sigma N^2/2)k\Delta + \sum_{i=0}^{k-1} \sigma N \Delta B_i
\]

for any integer \( k \geq 1 \). Thus, for any \( p > 0 \) we have

\[
\mathbb{E}[e^{pY_k}] \leq \mathbb{E}\left[ \exp\left( pY_0 + p(\eta + \sigma N^2/2)T + p\sum_{i=0}^{k-1} \sigma N \Delta B_i \right) \right]
\leq \exp\left( pY_0 + p(\eta + \sigma N^2/2)T + \frac{p^2\sigma^2 N^2}{2}T \right).
\]

The proof is complete. \( \square \)

Lemma 3.2. For any \( p > 0 \), the TEM scheme defined by (3.1) has the property that

\[
\sup_{\Delta \in (0,1)} \sup_{0 \leq k \leq [T/\Delta]} \mathbb{E}[e^{-pY_k}] \leq CT
\]

for any \( \Delta \in (0,1] \) and \( T > 0 \).

Proof. Using the well-known Taylor formula we get

\[
e^{-\bar{Y}_{k+1}} \leq e^{-Y_k}(\bar{Y}_{k+1} - Y_k) + \frac{1}{2}e^{-Y_k}(\bar{Y}_{k+1} - Y_k)^2 + \frac{1}{6}e^{-Y_k}(\bar{Y}_{k+1} - Y_k)^3,
\]

where \( |\bar{Y}_{k+1} - Y_k| \leq C\Delta + (\mu + \gamma)e^{Y_k}\Delta + |\sigma N|\Delta B_k| \leq C\Delta^3 + |\sigma N|\Delta B_k|, \) and

\[
|\bar{Y}_{k+1} - Y_k|^3 = |F(Y_k)\Delta + \sigma N \Delta B_k|^3 \leq 4(|F(Y_k)|^3\Delta^3 + |\sigma N|^3|\Delta B_k|^3)
\leq 4(4\eta + 0.5\sigma^2 N^2|\Delta^3 + 4(\mu + \gamma)^3e^{3Y_k}\Delta^3 + |\sigma N|^3|\Delta B_k|^3)
\leq 4(4\eta + 0.5\sigma^2 N^2|\Delta^3 + 4K^3(\mu + \gamma)^3\Delta^{3/2} + |\sigma N|^3|\Delta B_k|^3). \quad (3.5)
\]

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Then using the above inequality, we have

\[
e^{-Y_{k+1}} \leq e^{-Y_k} \left[ 1 - (Y_{k+1} - Y_k) + \frac{1}{2}(Y_{k+1} - Y_k)^2 + C \exp \left( C\Delta^\frac{3}{2} + |s\Delta\Delta B_k| \right) \left( \Delta^\frac{3}{2} + |\Delta B_k|^3 \right) \right]
\]

\[
\leq e^{-Y_k} \left[ 1 - (Y_{k+1} - Y_k) + \frac{1}{2}(Y_{k+1} - Y_k)^2 + C \exp \left( (s\Delta\Delta B_k) \left( \Delta^\frac{3}{2} + |\Delta B_k|^3 \right) \right) \right]
\]

\[
e^{-Y_k} \left[ 1 - F(Y_k)\Delta - s\Delta N\Delta B_k + \frac{1}{2}(F(Y_k)\Delta + s\Delta N\Delta B_k)^2 + U_k \right]
\]

\[
\leq e^{-Y_k} \left[ 1 - F(Y_k)\Delta - s\Delta N\Delta B_k + |F(Y_k)|^2\Delta^2 + |s\Delta|^2\Delta B_k|^2 + U_k \right]
\]

\[
\leq e^{-Y_k} \left[ 1 + ((\mu + \gamma)e^{Y_k} - \eta)\Delta - s\Delta N\Delta B_k + C\Delta^2 + K^2(\mu + \gamma)\Delta + |s\Delta|^2|\Delta B_k|^2 + U_k \right],
\]

where

\[
U_k = C \exp \left( (s\Delta\Delta B_k) \left( \Delta^\frac{3}{2} + |\Delta B_k|^3 \right) \right),
\]

by Lemma 2.3 one observes that

\[
\mathbb{E}[U_k^m | \mathcal{F}_{t_k}] \leq C\Delta^{\frac{3m}{2}},
\]

(3.6)

Then

\[
(1 + e^{-Y_{k+1}})^p \leq (1 + e^{-Y_k})^p(1 + \zeta_k)^p,
\]

(3.7)

where

\[
\zeta_k = \frac{e^{-Y_k}}{1 + e^{-Y_k}} \left[ (\mu + \gamma)e^{Y_k}\Delta + (K^2(\mu + \gamma) - \eta)\Delta + C\Delta^2 + |s\Delta|^2|\Delta B_k|^2 - s\Delta N\Delta B_k + U_k \right],
\]

and we can see that \(\zeta_k > -1\). For the given constant \(p > 2\), choose an integer \(m\) such that \(2m < p < 2(m + 1)\). It follows from [22, Lemma 3.3] and (3.7) that

\[
\mathbb{E}\left[ (1 + e^{-Z_{k+1}})^p | \mathcal{F}_{t_k} \right] \leq (1 + e^{-Y_k})^p \left\{ 1 + p\mathbb{E}[\zeta_k | \mathcal{F}_{t_k}] + \frac{p(p - 1)}{2}\mathbb{E}[\zeta_k^2 | \mathcal{F}_{t_k}] + \mathbb{E}[P_m(\zeta_k)\zeta_k^2 | \mathcal{F}_{t_k}] \right\},
\]

(3.8)

where \(P_m(x)\) represents a \(m\)th-order polynomial of \(x\) with coefficients depending only on \(p\), and \(m\) is an integer. Noticing that the increment \(\Delta B_k\) is independent of \(\mathcal{F}_{t_k}\), we derive that

\[
\mathbb{E}[\langle \Delta B_k \rangle^{2j-1} | \mathcal{F}_{t_k}] = 0, \quad \mathbb{E}[\langle \Delta B_k \rangle^{2j} | \mathcal{F}_{t_k}] = \frac{(2j)!}{2^j j!} \Delta^j, \quad j = 1, 2, \ldots,
\]

(3.9)

and using (3.6), we compute

\[
\mathbb{E}[\zeta_k | \mathcal{F}_{t_k}] = \frac{e^{-Y_k}}{1 + e^{-Y_k}} \left[ (\mu + \gamma)e^{Y_k}\Delta + (K^2(\mu + \gamma) + |s\Delta|^2 - \eta)\Delta + C\Delta^2 + \mathbb{E}[U_k | \mathcal{F}_{t_k}] \right]
\]

\[
\leq \frac{e^{-Y_k}}{1 + e^{-Y_k}} \left[ (\mu + \gamma)e^{Y_k}\Delta + C\Delta \right] \leq C\Delta,
\]

(3.10)
and
\[
E[\zeta_k^3 | F_k] = \frac{e^{-2Y_k}}{(1 + e^{-Y_k})^3} E\left[ \left( (\mu + \gamma)e^{Y_k} \Delta + (K^2(\mu + \gamma) - \eta) \Delta + C \Delta^2 \right.ight.
\left. + |\sigma N|^2(\Delta B_k)^2 - \sigma N \Delta B_k + U_k \right)^2 | F_k \right]
\leq \frac{e^{-2Y_k}}{(1 + e^{-Y_k})^3} \left( 3(\mu + \gamma)^2 e^{2Y_k} \Delta^2 + C \Delta^2 + 3|\sigma N|^2 \Delta + E[U_k^2 | F_k] \right)
\leq \frac{e^{-2Y_k}}{(1 + e^{-Y_k})^3} \left( 3(\mu + \gamma)^2 e^{2Y_k} \Delta^2 + C \Delta \right) \leq C \Delta.
\]

(3.11)

To estimate \(E[P_m(\zeta_k)\zeta_k^3 | F_k]\), we begin with \(E[\zeta_k^3 | F_k]\). Using (3.6) and (3.9) we obtain
\[
E[\zeta_k^3 | F_k] = \frac{e^{-2Y_k}}{(1 + e^{-Y_k})^3} E\left[ \left( (\mu + \gamma)e^{Y_k} \Delta + (K^2(\mu + \gamma) - \eta) \Delta + C \Delta^2 \right.ight.
\left. + |\sigma N|^2(\Delta B_k)^2 - \sigma N \Delta B_k + U_k \right)^3 | F_k \right]
\leq \frac{e^{-3Y_k}}{(1 + e^{-Y_k})^3} \left( (\mu + \gamma)^3 e^{3Y_k} \Delta^3 + C \Delta^3 + E[U_k^3 | F_k] \right) \leq C \Delta^\frac{3}{2}.
\]

(3.12)

On the other hand, we can use the same method to derive that
\[
E[\zeta_k^3 | F_k] = \frac{e^{-3Y_k}}{(1 + e^{-Y_k})^3} E\left[ \left( (\mu + \gamma)e^{Y_k} \Delta + (K^2(\mu + \gamma) - \eta) \Delta + C \Delta^2 \right.ight.
\left. + |\sigma N|^2(\Delta B_k)^2 - \sigma N \Delta B_k + U_k \right)^3 | F_k \right]
\geq - \frac{e^{-3Y_k}}{(1 + e^{-Y_k})^3} \left( (\mu + \gamma)^3 e^{3Y_k} \Delta^3 + C \Delta^3 + E[U_k^3 | F_k] \right) \geq -C \Delta^\frac{3}{2}.
\]

(3.13)

Thus, both of the above inequality imply \(E[a_i a_0 \zeta_k^3 | F_k] \leq o(\Delta)\) for any constant \(a_0\), where \(a_i\) represents the coefficient of \(\zeta_k^i\) term in polynomial \(P_m(\zeta_k)\). We can also show that
\[
E[a_i a_0 \zeta_k^{3+i} | F_k] \leq o(\Delta)
\]
for any \(i \geq 1\). These implies
\[
E[P_m(\zeta_k)\zeta_k^3 | F_k] \leq o(\Delta).
\]

(3.14)

Combining (3.8), (3.10), (3.11) and (3.14), we obtain that
\[
E \left[ (1 + e^{-Y_{k+1}})^p | F_k \right] \leq (1 + e^{-Y_k})^p (1 + C \Delta)
\]

(3.15)

for any integer \(0 \leq k \leq \lfloor T/\Delta \rfloor\). Obviously,
\[
E \left[ (1 + e^{-Y_{k+1}})^p \right] \leq (1 + C \Delta) E \left[ (1 + e^{-Y_k})^p \right].
\]

(3.16)
Define $\Omega_k = \{\omega | \bar{Y}_k \leq \log(K\Delta^{-\frac{1}{2}})\}$ and using the Chebyshev’s inequality, we can see that
\[
\mathbb{E}\left[(1 + e^{-Y_k})^p\right] = \mathbb{E}\left[(1 + e^{-Y_k})^p I_{\Omega_k}\right] + \mathbb{E}\left[(1 + e^{-Y_k})^p I_\mathbb{R}\right]
\leq \mathbb{E}\left[(1 + e^{-Y_k})^p\right] + 2^p\mathbb{P}\{\bar{Y}_k > \log(K\Delta^{-\frac{1}{2}})\}
\leq (1 + C\Delta)\mathbb{E}\left[(1 + e^{-Y_{k-1}})^p\right] + 2^p \frac{Ee^{2Y_k}}{K^2\Delta^T}.
\] (3.17)

It follows from the result of Lemma 3.1 that
\[
\mathbb{E}\left[(1 + e^{-Y_k})^p\right] \leq (1 + C\Delta)\mathbb{E}\left[(1 + e^{-Y_{k-1}})^p\right] + C\Delta
\leq (1 + C\Delta)^k (1 + e^{-Y_0})^p + C\Delta \sum_{i=0}^{k-1} (1 + C\Delta)^i
\leq e^{Ck\Delta} (1 + e^{-Y_0})^p + Ck\Delta e^{Ck\Delta} \leq e^{CT} \left[(1 + e^{-Y_0})^p + CT\right].
\] (3.18)

Therefore the desired result follows. \[ \square \]

Define $\tau_R = \inf\{k\Delta \geq 0 : y(t_k) \geq \log R\}$, $\rho_\Delta = \inf\{k\Delta \geq 0 : \bar{Y}_\Delta(t_k) \geq \log(K\Delta^{-\frac{1}{2}})\}$. By Lemma 2.2 we have
\[
R^p\mathbb{P}\{\tau_R \leq T\} \leq \mathbb{E}\left[e^{p\bar{Y}_\Delta(T \wedge \tau_R)} I_{\{\tau_R \leq T\}}\right] + \mathbb{E}\left[e^{p\bar{Y}_\Delta(T \wedge \tau_R)} I_{\{\tau_R > T\}}\right] = \mathbb{E}\left[e^{p\bar{Y}_\tau(T \wedge \tau)}\right] \leq C_T.
\]

It is easy to see that
\[
\mathbb{P}\{\tau_R \leq T\} \leq C_T R^{-p}, \quad \forall \ p > 0,
\] (3.19)

where $C$ is a positive constant independent of $R$. Moreover, by the virtue of Lemma 3.1, we have
\[
\mathbb{E}\left[\exp\left(p\log(K\Delta^{-\frac{1}{2}})\right) I_{\{\rho_\Delta \leq T\}}\right] \leq \mathbb{E}\left[\exp\left(p\bar{Y}_\Delta(T \wedge \rho_\Delta)\right)\right] = \mathbb{E}\left[\exp\left(p\bar{Y}_{\tau(\Delta)}\right)\right]
\leq \sup_{0 \leq k \leq \lfloor T/\Delta \rfloor} \mathbb{E}\left[\exp\left(p\bar{Y}_k\right)\right] \leq C_T,
\]
implies that
\[
\mathbb{P}\{\rho_\Delta \leq T\} \leq C_T K^{-p} k^{p/2}, \quad \forall \ p > 0,
\] (3.20)

where $C$ is a positive constant independent of $\Delta$.

**Theorem 3.3.** For any $p > 0$, the TEM scheme defined by (3.1) has the property that
\[
\sup_{0 \leq k \leq \lfloor T/\Delta \rfloor} \mathbb{E}\left[|Y_k - y(t_k)|^p\right] \leq C_T \Delta^p
\]
for any $\Delta \in (0, 1)$ and $T > 0$. 

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Proof. Define $\bar{\theta}_\Delta = \tau_{K\Delta^{-1/2}} \wedge \rho_\Delta$, $\Omega_1 := \{ \omega : \bar{\theta}_\Delta > T \}$, $\bar{u}_k = Y_k - y(t_k)$, for any $k \Delta \in [0, T]$, where $\tau_L$ and $\rho_\Delta$ are defined by (3.19), and (3.20), respectively. For any $q > p$, using the Young inequality we obtain that
\[
\mathbb{E}|\bar{u}_k|^p \leq \mathbb{E}(|\bar{u}_k|^p1_{\Omega_1}) + \mathbb{E}(|\bar{u}_k|^p1_{\Omega_2}) \\
\leq \mathbb{E}(|\bar{u}_k|^p1_{\Omega_1}) + \frac{p \Delta^p}{q} \mathbb{E}(|\bar{u}_k|^q) + \frac{q - p}{q \Delta^p/(q - p)} \mathbb{P}(\Omega_1^c).
\]
(3.21)

It follows from the results of Lemmas 2.2, 3.1 and 3.2 that
\[
\frac{p \Delta^p}{q} \mathbb{E}(|\bar{u}_k|^q) \leq C T \Delta^p.
\]
(3.22)

It follows from (3.19), and (3.20) that
\[
\frac{q - p}{q \Delta^p/(q - p)} \mathbb{P}(\Omega_1^c) \leq \frac{q - p}{q \Delta^p/(q - p)} \left( \mathbb{P}\{\tau_{K\Delta^{-1/2}} \leq T \} + \mathbb{P}\{\rho_\Delta \leq T \} \right) \\
\leq \frac{2(q - p)}{q \Delta^p/(q - p)} \frac{C_T}{K^\delta \Delta^{-\frac{p}{2}}} \leq C_T \Delta^\frac{q}{2} \Delta^{-\frac{p}{2}} \leq C_T \Delta^p,
\]
(3.23)

where $\delta \geq 2pq/(q - p)$. Inserting (3.22), (3.23) and (2.5) into (3.21) yields
\[
\mathbb{E}|\bar{u}_k|^p \leq \mathbb{E}(|\bar{u}_k|^p1_{\Omega_1}) + C_T \Delta^p \leq \sup_{0 \leq k \leq \lceil T/\Delta \rceil} \mathbb{E}\left[|Y_k - y(t_k)|^p\right] + C_T \Delta^p
\]
\[
= \sup_{0 \leq k \leq \lceil T/\Delta \rceil} \mathbb{E}\left[|Y^E_k - y(t_k)|^p\right] + C_T \Delta^p \leq C_T \Delta^p.
\]
(3.24)

The proof is complete. □

Theorem 3.4. For any $p > 0$ there exists a constant $C > 0$ such that
\[
\sup_{k=0, \ldots, \lfloor T/\Delta \rfloor} \mathbb{E}\left[|I(t_k) - I_{\Delta_k}|^p\right] \leq C_T \Delta^p
\]
for any $\Delta \in (0, 1]$ and $T > 0$.

Proof. By the Lamperti-transformation (1.5) and (3.2), we have
\[
\sup_{k=0, \ldots, \lfloor T/\Delta \rfloor} \mathbb{E}\left[|I(t_k) - I_{\Delta_k}|^p\right] = N \sup_{k=0, \ldots, \lfloor T/\Delta \rfloor} \mathbb{E}\left[\left|\frac{e^{y(t_k)}}{1 + e^{y(t_k)}} - \frac{e^{Y_k}}{1 + e^{Y_k}}\right|^p\right] \\
\leq N \sup_{k=0, \ldots, \lfloor T/\Delta \rfloor} \mathbb{E}\left[|y(t_k) - Y_k|^p\right].
\]

Thus, by applying Theorem 3.3, we infer that
\[
\mathbb{E}\left[\sup_{k=0, \ldots, \lfloor T/\Delta \rfloor} |I(t_k) - I_{\Delta_k}|^p\right] \leq N \sup_{k=0, \ldots, \lfloor T/\Delta \rfloor} \mathbb{E}\left[|y(t_k) - Y_k|^p\right] \leq C_T \Delta^p
\]
(3.25)

for any $\Delta \in (0, 1]$. The proof is complete. □
4. Extinction

In the study of population systems, extinction is one of the most important issues, the easily implementable scheme preserving the underlying extinction is desired eagerly. In this section we cite the extinction criterion of the exact solution for SIS epidemic models (1.2) (see, e.g., [2]) and go a further step to show the logarithmic TEM method keeping extinction well. To begin this section, we cite an extinction of stochastic SIS models, please see details in [2].

Theorem 4.1 ([2, Theorem 4.1]). If

\[ R_0^S := R_0^D - \frac{\sigma^2 N^2}{2(\mu + \gamma)} = \frac{\beta N}{\mu + \gamma} - \frac{\sigma^2 N^2}{2(\mu + \gamma)} < 1 \quad \text{and} \quad \sigma^2 \leq \frac{\beta}{N}, \]  

then for any given initial value \( I_0 \in (0, N) \), the solution of the SDE (1.2) obeys

\[ \limsup_{t \to \infty} \frac{1}{t} \log(I(t)) \leq \beta N - \mu - \gamma - 0.5 \sigma^2 N^2 < 0 \quad \text{a.s.;} \]  

namely, \( I(t) \) tends to zero exponentially a.s. In other words, the disease dies out with probability one.

Theorem 4.2. Under the condition of Theorem 4.1, there is a constant \( \Delta \in (0, \Delta^*) \) such that the scheme (3.1) has the property that

\[ \limsup_{k \Delta \to \infty} \frac{1}{k\Delta} \log(I_k^\Delta) \leq \beta N - \mu - \gamma - 0.5 \sigma^2 N^2 + h(\Delta) < 0 \quad \text{a.s.;} \]

namely, \( I_k^\Delta \) tends to zero exponentially a.s. In other words, the disease dies out with probability one, where

\[ h(\Delta) := 6|\eta + 0.5 \sigma^2 N^2| \Delta^2 + 6K^3(\mu + \gamma)^3 + 4|\sigma N|^3 \Delta^2, \]  

and \( \Delta^* \) is the positive solution of the equation (in \( \Delta \))

\[ h(\Delta) = 0.5 \sigma^2 N^2 + \mu + \gamma - \beta N. \]

Proof. Clear, we can see that \( Y_k \leq \log(K \Delta^{-\frac{1}{2}}) \) and \( \bar{Y}_k \leq \bar{Y}_k \). Direct calculation leads to

\[ I_k^\Delta - \bar{I}_k = \frac{N}{1 + e^{Y_k}} - \frac{N}{1 + e^{\bar{Y}_k}} = \frac{N(e^{Y_k} - e^{\bar{Y}_k})}{(1 + e^{Y_k})(1 + e^{\bar{Y}_k})} \leq 0, \]

implies that \( I_k \leq \bar{I}_k \). Here and below we always denote

\[ f(x) := N - \frac{N}{1 + e^x} = \frac{Ne^x}{1 + e^x}, \quad \forall \ x \in \mathbb{R}, \]

we derive that

\[ \begin{cases} f'(x) = \frac{Ne^x}{1 + e^x} \frac{1}{1 + e^x} = \frac{f(x)}{1 + e^x}, & f''(x) = \frac{Ne^x(1 - e^x)}{(1 + e^x)^3} = \frac{f(x)(1 - e^x)}{(1 + e^x)^2}, \\ f'''(x) = \frac{Ne^x}{(1 + e^x)^4} \left[ (1 - e^x)^2 - 2e^x \right] = \frac{f(x)}{(1 + e^x)^2} \left[ (1 - e^x)^2 - 2e^x \right]. \end{cases} \]  

(4.5)
To proceed, we define

\[ V(x) := \log f(x) = \log \frac{N e^x}{1 + e^x}, \quad \forall \ x \in \mathbb{R}. \]

It is easy to see that

\[
\left\{
\begin{array}{l}
V'(x) = \frac{f'(x)}{f(x)} = \frac{1}{1 + e^x}, \\
V''(x) = \left[ \frac{f'(x)}{f(x)} \right]' = \frac{f''(x)}{f(x)} - \left( \frac{f'(x)}{f(x)} \right)^2 = \frac{1 - e^x}{(1 + e^x)^2} - \frac{1}{(1 + e^x)^2} = \frac{-e^x}{(1 + e^x)^2},
\end{array}
\right.
\tag{4.6}
\]

and for any \( x \in \mathbb{R} \),

\[
|V''(x)| \leq \frac{1}{(1 + e^x)^2} + \frac{1}{(1 + e^x)^2} \leq 2.
\tag{4.7}
\]

Then, by the virtue of the Taylor formula with Lagrange remainder term, we have

\[
V(\bar{Y}_{k+1}) = V(Y_k) + V'(Y_k)(\bar{Y}_{k+1} - Y_k) + \frac{V''(Y_k)}{2} (\bar{Y}_{k+1} - Y_k)^2 \\
+ \frac{V'''(Y_k + \theta(\bar{Y}_{k+1} - Y_k))}{3!} (\bar{Y}_{k+1} - Y_k)^3
= V(Y_k) + \frac{\bar{Y}_{k+1} - Y_k}{1 + e^{Y_k}} + \frac{-e^{Y_k}(\bar{Y}_{k+1} - Y_k)^2}{2(1 + e^{Y_k})^2} + \frac{V'''(Y_k + \theta(\bar{Y}_{k+1} - Y_k))}{3!}(\bar{Y}_{k+1} - Y_k)^3,
\tag{4.8}
\]

where \( \theta \in (0, 1) \). By (4.7), we can see that

\[
V(\bar{Y}_{k+1}) \leq V(Y_k) + (1 + e^{Y_k})^{-1}(\bar{Y}_{k+1} - Y_k) + \frac{-e^{Y_k}}{2(1 + e^{Y_k})^2}(\bar{Y}_{k+1} - Y_k)^2 + \frac{1}{3}\bar{Y}_{k+1} - Y_k|^3.
\tag{4.9}
\]

It follows from Lemma 2.3 that \( \mathbb{E}|\triangle B_k|^3 = (2\triangle)^{\frac{3}{2}}/\sqrt{\pi} \), and by (3.5), one observes

\[
\frac{1}{3}\bar{Y}_{k+1} - Y_k|^3 \leq h(\triangle)\triangle + M_k^{(1)},
\]

where

\[
M_k^{(1)} = 2|\sigma N|^3\left( |\triangle B_k|^3 - \frac{(2\triangle)^{\frac{3}{2}}}{\sqrt{\pi}} \right).
\]

Inserting the above inequality into (4.9) we have

\[
V(\bar{Y}_{k+1}) \leq V(Y_k) + (1 + e^{Y_k})^{-1}(\bar{Y}_{k+1} - Y_k) + \frac{-e^{Y_k}}{2(1 + e^{Y_k})^2}(\bar{Y}_{k+1} - Y_k)^2 + h(\triangle)\triangle + M_k^{(1)}.
\tag{4.10}
\]

The second and third terms at the right end of the estimation inequality, we have

\[
(1 + e^{Y_k})^{-1}(\bar{Y}_{k+1} - Y_k) = (1 + e^{Y_k})^{-1}\left[ (\eta - (\mu + \gamma)e^{Y_k} + \frac{\sigma^2 N^2}{2} - \frac{\sigma^2 N^2}{1 + e^{Y_k}})\triangle + \sigma N\triangle B_k \right]
\]

\[
= (1 + e^{Y_k})^{-1}(\eta - (\mu + \gamma)e^{Y_k})\triangle + (1 + e^{Y_k})^{-1}\sigma N\triangle B_k
+ (1 + e^{Y_k})^{-1}\left( \frac{\sigma^2 N^2}{2} - \frac{\sigma^2 N^2}{1 + e^{Y_k}} \right)\triangle,
\tag{4.11}
\]

\[ 12 \]
and

\[(\bar{Y}_{k+1} - Y_k)^2 = \left[ (\eta - (\mu + \gamma)e^{Y_k} + \frac{\sigma^2N^2}{2} - \frac{\sigma^2N^2}{1 + e^{Y_k}}) \Delta + \sigma N \Delta B_k \right]^2 \]

\[= \left( \eta - (\mu + \gamma)e^{Y_k} + \frac{\sigma^2N^2}{2} - \frac{\sigma^2N^2}{1 + e^{Y_k}} \right)^2 \Delta^2 + \sigma^2N^2(\Delta B_k)^2 \Delta + \sigma^2N^2 \Delta \]

\[+ 2\sigma N \Delta \left( \eta - (\mu + \gamma)e^{Y_k} + \frac{\sigma^2N^2}{2} - \frac{\sigma^2N^2}{1 + e^{Y_k}} \right) \Delta B_k \]

\[= \sigma^2N^2 \Delta + \left( \eta - (\mu + \gamma)e^{Y_k} + \frac{\sigma^2N^2}{2} - \frac{\sigma^2N^2}{1 + e^{Y_k}} \right)^2 \Delta^2 + M_k^{(2)}, \]

where

\[M_k^{(2)} := \sigma^2N^2(\Delta B_k)^2 \Delta + 2\sigma N \Delta \left( \eta - (\mu + \gamma)e^{Y_k} + \frac{\sigma^2N^2}{2} - \frac{\sigma^2N^2}{1 + e^{Y_k}} \right) \Delta B_k. \]

Then

\[\frac{-e^{Y_k}}{2(1 + e^{Y_k})^2} (\bar{Y}_{k+1} - Y_k)^2 = \frac{-e^{Y_k}}{2(1 + e^{Y_k})^2} \left[ (\eta - (\mu + \gamma)e^{Y_k} + \frac{\sigma^2N^2}{2} - \frac{\sigma^2N^2}{1 + e^{Y_k}}) \Delta + \sigma N \Delta B_k \right]^2 \]

\[= - \frac{\sigma^2N^2e^{Y_k}}{2(1 + e^{Y_k})^2} \Delta - \frac{e^{Y_k}}{2(1 + e^{Y_k})^2} M_k^{(2)} \]

\[- \frac{e^{Y_k}}{2(1 + e^{Y_k})^2} \left( \eta - (\mu + \gamma)e^{Y_k} + \frac{\sigma^2N^2}{2} - \frac{\sigma^2N^2}{1 + e^{Y_k}} \right)^2 \Delta^2 + h(\Delta) \Delta + M_k^{(1)} \]

\[\leq \frac{-e^{Y_k}}{2(1 + e^{Y_k})^2} \Delta + h(\Delta) \Delta + M_k, \]

Substituting (4.11) and (4.12) into (4.10) yields

\[V(\bar{Y}_{k+1}) \leq V(Y_k) + (1 + e^{Y_k})^{-1}(\eta - (\mu + \gamma)e^{Y_k}) \Delta + (1 + e^{Y_k})^{-1}\sigma N \Delta B_k \]

\[+ (1 + e^{Y_k})^{-1} \left( \frac{\sigma^2N^2}{2} - \frac{\sigma^2N^2}{1 + e^{Y_k}} \right) \Delta - \frac{\sigma^2N^2e^{Y_k}}{2(1 + e^{Y_k})^2} \Delta - \frac{\sigma^2N^2e^{Y_k}}{2(1 + e^{Y_k})^2} M_k^{(2)} \]

\[- \frac{e^{Y_k}}{2(1 + e^{Y_k})^2} \left( \eta - (\mu + \gamma)e^{Y_k} + \frac{\sigma^2N^2}{2} - \frac{\sigma^2N^2}{1 + e^{Y_k}} \right)^2 \Delta^2 + h(\Delta) \Delta + M_k^{(1)} \]

\[\leq V(Y_k) + (1 + e^{Y_k})^{-1}(\eta - (\mu + \gamma)e^{Y_k}) \Delta + (1 + e^{Y_k})^{-1} \left( \frac{\sigma^2N^2}{2} - \frac{\sigma^2N^2}{1 + e^{Y_k}} \right) \Delta \]

\[- \frac{\sigma^2N^2e^{Y_k}}{2(1 + e^{Y_k})^2} \Delta + h(\Delta) \Delta + M_k, \]

where

\[M_k = (1 + e^{Y_k})^{-1}\sigma N \Delta B_k - \frac{e^{Y_k}}{2(1 + e^{Y_k})^2} M_k^{(2)} + M_k^{(1)}. \]

We can see that

\[(1 + e^{Y_k})^{-1} \left( \frac{\sigma^2N^2}{2} - \frac{\sigma^2N^2}{1 + e^{Y_k}} \right) - \frac{\sigma^2N^2e^{Y_k}}{2(1 + e^{Y_k})^2} = \frac{\sigma^2N^2}{2(1 + e^{Y_k})^2} (1 + e^{Y_k} - 2) = - \frac{\sigma^2N^2}{2(1 + e^{Y_k})^2}, \]

and

\[(1 + e^{Y_k})^{-1}(\eta - (\mu + \gamma)e^{Y_k}) = (1 + e^{Y_k})^{-1}(\beta N - \mu - \gamma - (\mu + \gamma)e^{Y_k}) \]

\[= (1 + e^{Y_k})^{-1}(\beta N - \mu - \gamma - (\mu + \gamma)e^{Y_k}). \]
\[ (1 + e^{Y_k})^{-1}(\beta N - (\mu + \gamma)(1 + e^{Y_k})) = \frac{\beta N}{1 + e^{Y_k}} - \mu - \gamma. \]

Making use of the above inequality and (4.13) yields
\[ V(Y_{k+1}) \leq V(Y_k) + \left( \frac{\beta N}{1 + e^{Y_k}} - \mu - \gamma - \frac{\sigma^2 N^2}{2(1 + e^{Y_k})^2} \right) \Delta + h(\Delta) \Delta + M_k. \quad (4.14) \]

Due to (3.2) and definition of \( V \), we have
\[ V(\bar{Y}_{k+1}) = \log(I_{k+1}^\Delta), \quad V(Y_k) = \log(I_k^\Delta), \quad N - I_k^\Delta = \frac{N}{1 + e^{Y_k}}, \]
and hence
\[ V(Y_{k+1}) \leq V(\bar{Y}_{k+1}), \]
\[ \frac{\beta N}{1 + e^{Y_k}} - \mu - \gamma = \beta N - \beta I_k^\Delta - \mu - \gamma, \quad - \frac{\sigma^2 N^2}{2(1 + e^{Y_k})^2} = -0.5\sigma^2(N - I_k^\Delta)^2. \]

Then
\[ \log(I_{k+1}^\Delta) \leq \log(I_k^\Delta) + \left[ \beta N - \mu - \gamma - \beta I_k^\Delta - 0.5\sigma^2(N - I_k^\Delta)^2 \right] \Delta + h(\Delta) \Delta + M_k \]
\[ = \log(I_k^\Delta) + g(I_k^\Delta) \Delta + h(\Delta) \Delta + M_k, \quad (4.15) \]
for any integer \( k \geq 0 \), where \( g : \mathbb{R} \to \mathbb{R} \) is defined by
\[ g(x) = \beta N - \mu - \gamma - \beta x - 0.5\sigma^2(N - x)^2. \quad (4.16) \]

However, under condition (4.1), we have
\[ g(I_k^\Delta) \leq \beta N - \mu - \gamma - 0.5\sigma^2 N^2 \]
for \( I_k^\Delta \in (0, N) \). It then follows from (4.15) that
\[ \log(I_{k+1}^\Delta) \leq \log(I_k^\Delta) + (\beta N - \mu - \gamma - 0.5\sigma^2 N^2) \Delta + h(\Delta) \Delta + M_k. \]

Repeating this procedure arrives at
\[ \log(I_k^\Delta) \leq \log(I_0) + (\beta N - \mu - \gamma - 0.5\sigma^2 N^2) k \Delta + h(\Delta) k \Delta + \sum_{i=0}^{k-1} M_i. \]

This implies
\[ \limsup_{k \Delta \to +\infty} \frac{1}{k \Delta} \log(I_k^\Delta) \leq \beta N - \mu - \gamma - 0.5\sigma^2 N^2 + h(\Delta) + \limsup_{k \Delta \to +\infty} \frac{1}{k \Delta} \sum_{i=0}^{k-1} M_i \quad \text{a.s.} \quad (4.17) \]

However, by the large number theorem for martingales (see, e.g., [23]), we have
\[ \limsup_{k \Delta \to +\infty} \frac{1}{k \Delta} \sum_{i=0}^{k-1} M_i = 0 \quad \text{a.s.} \]

We therefore obtain the desired assertion (4.20) from (4.17). \( \Box \)

In Theorem 4.1 we require the noise intensity \( \sigma^2 \leq \beta/N \). The following theorem covers the case when \( \sigma^2 > \beta/N \).
Theorem 4.3 ([2, Theorem 4.3]). If
\[
\sigma^2 > \frac{\beta}{N} \sqrt{\frac{\beta^2}{2(\mu + \gamma)}}, \tag{4.18}
\]
then for any given initial value \(I_0 \in (0, N)\), the solution of the SDE SIS model (1.2) obeys
\[
\limsup_{t \to \infty} \frac{1}{t} \log(I(t)) \leq -\mu - \gamma + \frac{\beta^2}{2\sigma^2} < 0 \quad \text{a.s.}; \tag{4.19}
\]
namely, \(I(t)\) tends to zero exponentially a.s. In other words, the disease dies out with probability one.

Note that condition (4.18) implies that \(R_S^* \leq 1\).

Theorem 4.4. Under the condition of Theorem 4.3, there is a constant \(\triangle \in (0, \Delta^{**})\) such that the scheme (3.1) has the property that
\[
\limsup_{k \Delta \to +\infty} \frac{1}{k \Delta} \log(I_k^\Delta) \leq -\mu - \gamma + \frac{\beta^2}{2\sigma^2} + h(\Delta) < 0 \quad \text{a.s.}; \tag{4.20}
\]
namely, \(I_k^\Delta\) tends to zero exponentially a.s. In other words, the disease dies out with probability one, where \(\Delta^{**}\) is the positive solution of the equation (in \(\triangle\))
\[
h(\Delta) = \mu + \gamma - \frac{\beta^2}{2\sigma^2},
\]
and \(h(\Delta)\) is given by (4.4).

Proof. We use the same way as in the proof of Theorem 4.2. By condition (4.18), it is easy to see that \(g(x) \leq g(\bar{x})\) for any \(x \in (0, N)\), where
\[
\bar{x} := \frac{\sigma^2 N - \beta}{\sigma^2} \in (0, N).
\]
Compute
\[
g(\bar{x}) = \beta N - \mu - \gamma - 0.5\sigma^2 N^2 + \frac{(\sigma^2 N - \beta)^2}{2\sigma^2} = -\mu - \gamma + \frac{\beta^2}{2\sigma^2} < 0.
\]
This implies, in the same way as in the proof of Theorem 4.2, that
\[
\limsup_{k \Delta \to +\infty} \frac{1}{k \Delta} \log(I_k^\Delta) \leq g(\bar{x}) + h(\Delta) \quad \text{a.s.},
\]
as required. The proof is hence complete.  \(\square\)
5. Examples and simulations

In order to illustrate the efficiency of the logarithmic TEM scheme we introduce some examples and simulations. We shall assume that the unit of time is one day and the population sizes are measured in units of 1 million, unless otherwise stated.

Example 5.1. Consider the stochastic SIS epidemic model (see, e.g., [2, Example 4.2])

\[
dI(t) = \left[5I(t) - 0.5I(t)^2\right]dt + 0.035I(t)(100 - I(t))dB(t).
\] (5.1)

Noting that

\[
R_0^S = \frac{\beta N}{\mu + \gamma} - \frac{\sigma^2 N^2}{2(\mu + \gamma)} = 1.111 - 0.136 < 1 \quad \text{and} \quad \sigma^2 = 0.001225 \leq \frac{\beta}{N} = 0.005,
\]

we can therefore conclude, by Theorem 4.1, that for any initial value \(I_0 \in (0, 100)\), \(I(t)\) will tend to zero exponentially with probability one.

To compare to the simulations of the classical EM method (see, [2, Example 4.2]), then all parameters are same as the classical EM method. The simulations in Figure 2 show the sample paths of the solution for \(t \in [0, 50]\) by the logarithmic TEM scheme. Both figures show clearly that the logarithmic TEM scheme (3.2) reproduces the dynamic properties of the underlying SDE (5.1).

Due to no closed-form of the solution, using scheme (3.2), we regard the better approximation with \(\Delta = 2^{-19}\) as the exact solution \(I(t)\) and compare it with the numerical solution \(I_k\) with \(\Delta = 2^{-9}, 2^{-10}, \ldots, 2^{-16}\). To compute the approximation error, we run \(M\) independent trajectories where \(I^{(j)}(t_k)\) and \(I_k^{(j)}\) represent the \(j\)th trajectories of the exact solution \(I(t)\) and the numerical solution \(I_k\) respectively. Thus

\[
\text{Error}(p) := \left( \mathbb{E} \left[ \sup_{k=0,\ldots,[T/\Delta]} |I(t_k) - I_k|^{p} \right] \right)^{\frac{1}{p}} \approx \left( \frac{1}{M} \sum_{j=1}^{M} \left[ \sup_{k=0,\ldots,[T/\Delta]} |I^{(j)}(t_k) - I_k^{(j)}|^{p} \right] \right)^{\frac{1}{p}}.
\]
To reduce the time of simulations without losing any necessary illustration, we only simulate the paths for \( t \in [0, 2] \), and plotting the \( \log_2 \)-approximation error \( (\mathbb{E}\left[ \sup_{k=0,\ldots,\lfloor 2/t \rfloor} |I(t_k) - I_k^{(5)}| \right] )^{\frac{1}{2}} \) with \( M = 1000 \). The blue solid line depicts \( \log_2 \)-error while the red dashed is a reference line of slope 1 in Figure 1.

**Example 5.2.** We keep the system parameters the same as in Example 5.1 but let \( \sigma = 0.08 \), so the stochastic SIS model (1.2) becomes (see, e.g., [2, Example 4.4])

\[
dI(t) = \left[ 5I(t) - 0.5I(t)^2 \right] dt + 0.08I(t)(100 - I(t))dB(t).  \tag{5.2}
\]

It is easy to verify that the system parameters obey condition (4.18). We can therefore conclude, by Theorem 4.3, that for any initial value \( I_0 \in (0, 100) \), \( I(t) \) will tend to zero exponentially with probability one.

The computer simulations shown in Figure 3 clearly support these results.

Figure 2: The approximation error of the exact solution and the numerical solution by the logarithmic EM scheme (2.3) as the function of step size \( \Delta \in \{2^{-9}, 2^{-10}, \ldots, 2^{-16}\} \), with initial values (left) \( I_0 = 90 \) and (right) \( I_0 = 1 \).

Figure 3: The approximation error of the exact solution and the numerical solution by the logarithmic EM scheme (2.3) as the function of step size \( \Delta \in \{2^{-9}, 2^{-10}, \ldots, 2^{-16}\} \), with initial values (left) \( I_0 = 1 \) and (right) \( I_0 = 90 \).
Example 5.3. Consider the stochastic SIS epidemic model (see, e.g., [2, Example 5.3])

$$dI(t) = [5I(t) - 0.5I(t)^2]dt + 0.01I(t)(100 - I(t))dB(t). \quad (5.3)$$

The simulation descriptions of this example are similar to Example 5.1, so we omit it here.

Remark 5.4. This paper aims to show strong convergence rate and extinction of logarithmic TEM method for nonlinear SDEs (1.2) under nonglobally monotone coefficients. However, we have not obtained the persistence and stationary distribution of logarithmic TEM method, which could be a new direction of our further work. Although we have not proved the persistence of logarithmic TEM method, it is easy to see from Figure 6 that without extra restrictions the numerical solutions of the logarithmic TEM method stay in step of persistence with the exact solutions.

References

[1] H. Hethcote, J. Yorke, Gonorrhea Transmission Dynamics and Control, Springer, Berlin, 2014.
Figure 6: The approximation error of the exact solution and the numerical solution by the logarithmic EM scheme (2.3) as the function of step size $\Delta \in \{2^{-9}, 2^{-10}, \ldots , 2^{-16}\}$, with initial values (left) $I_0 = 90$ and (right) $I_0 = 1$.

[2] A. Gray, D. Greenhalgh, L. Hu, X. Mao, J. Pan, A stochastic differential equation SIS epidemic model, SIAM J. Appl. Math. 71 (2011) 876–902.

[3] M. Hutzenthaler, A. Jentzen, P. E. Kloeden, Strong and weak divergence in finite time of Euler’s method for stochastic differential equations with non-globally Lipschitz continuous coefficients, Proc. R. Soc. Lond. Ser. A Math. Phys. Eng. Sci. 467 (2011) 1563–1576.

[4] L. Szpruch, X. Mao, D. Higham, J. Pan, Numerical simulation of a strongly nonlinear Ait-Sahalia type interest rate model, BIT Numer. Math. 51 (2011) 405–425.

[5] X. Li, X. Mao, H. Yang, Strong convergence and asymptotic stability of explicit numerical schemes for nonlinear stochastic differential equations, Math. Comput. 90 (2021) 2827–2872.

[6] X. Mao, L. Szpruch, Strong convergence and stability of implicit numerical methods for stochastic differential equations with non-globally Lipschitz continuous coefficients, J. Comput. Appl. Math. 238 (2013) 14–28.

[7] M. Tretyakov, Z. Zhang, A fundamental mean-square convergence theorem for SDEs with locally Lipschitz coefficients and its applications, SIAM J. Numer. Anal. 51 (2013) 3135–3162.

[8] M. Hutzenthaler, A. Jentzen, P. E. Kloeden, Strong convergence of an explicit numerical method for SDEs with nonglobally Lipschitz continuous coefficients, Ann. Appl. Probab. 22 (4) (2012) 1611–1641.

[9] M. Hutzenthaler, A. Jentzen, Numerical approximations of stochastic differential equations with non-globally Lipschitz continuous coefficients, Mem. Amer. Math. Soc. 236 (1112) (2015) 99 pp.
[10] S. Sabanis, A note on tamed Euler approximations, Electron. Commun. Probab. 18 (13) (2013) 1–10.

[11] S. Sabanis, Euler approximations with varying coefficients: the case of superlinearly growing diffusion coefficients, Ann. Appl. Probab. 26 (4) (2016) 2083–2105.

[12] X. Wang, S. Gan, The tamed Milstein method for commutative stochastic differential equations with non-globally Lipschitz continuous coefficients, J. Difference Equ. Appl. 19 (3) (2013) 466–490.

[13] W. Liu, X. Mao, Strong convergence of the stopped Euler-Maruyama method for nonlinear stochastic differential equations, Appl. Math. Comput. 223 (4) (2013) 389–400.

[14] X. Mao, The truncated Euler-Maruyama method for stochastic differential equations, J. Comput. Appl. Math. 290 (2015) 370–384.

[15] X. Li, H. Yang, Explicit numerical approximation for logistic models with regime switching in finite and infinite horizons, (2021) arXiv:2106.03540.

[16] Y. Yi, Y. Hu, J. Zhao, Positivity preserving logarithmic Euler-Maruyama type scheme for stochastic differential equations, Commun Nonlinear Sci Numer Simulat 101 (2021) 105895.

[17] X. Mao, F. Wei, T. Wiriyakraikul, Positivity preserving truncated Euler-Maruyama Method for stochastic Lotka-Volterra competition model, J. Comput. Appl. Math. 394 (2021) 113566.

[18] L. Chen, S. Gan, X. Wang, First order strong convergence of an explicit scheme for the stochastic SIS epidemic model, J. Comput. Appl. Math. 392 (2021) 113482.

[19] H. Yang, J. Huang, First order strong convergence of positivity preserving logarithmic Euler-Maruyama method for the stochastic SIS epidemic model, Applied Mathematics Letters 121 (2021) 107451.

[20] A. Alfonsi, On the discretization schemes for the CIR (and Bessel squared) processes, Monte Carlo Methods Appl 11 (2005) 355–384.

[21] A. Neuenkirch, L. Szpruch, First order strong approximations of scalar SDEs defined in a domain, Numer. Math. 128 (2014) 103–136.

[22] H. Yang, X. Li, Explicit approximations for nonlinear switching diffusion systems in finite and infinite horizons, J. Differential Equations 265 (2018) 2921–2967.

[23] X. Mao, Stochastic Differential Equations and Applications 2nd edition, Horwood, Chichester, UK, 2007.