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Parental social class and GCSE attainment: Re-reading the role of ‘cultural capital’

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\textbf{ABSTRACT}
This paper examines the roles of parental social class and cultural capital in inequalities in English school qualifications. The analytical focus is the General Certificate of Secondary Education (GCSE). Integral to Bourdieu’s theory of cultural reproduction is the conception that inequalities in cultural capital explain the unequal scholastic achievements of pupils from different social class backgrounds. This paper is a novel investigation using data from \textit{Understanding Society} and linked administrative education records from the National Pupil Database. The central empirical findings do not support the Bourdieuian position, and there is no evidence that the substantial parental social class inequalities that are observed in school GCSE outcomes can be explained by inequalities in cultural capital. Engagement in reading related activities are mildly influential, but engagement in highbrow cultural activities are not influential. This is an important finding as the concept of cultural capital has become more prominent in Government education policy.

\textbf{Introduction}

The qualifications that British young people gain at school are consequential and affect their futures in both education and employment (Babb 2005, Banks et al. 1992, Furlong and Cartmel 2007, Jones, Joyce, and Thomas 2003, Raffe et al. 1998, Payne 2000, Sammons et al. 2014). Studies of the relationship between parental social class and young people’s outcomes in school qualifications report the overall finding that pupils from families in more advantaged social classes, on average, have more favourable outcomes (Connolly 2006, Demack, Drew, and Grimsley 2000, Drew, Gray, and Sime 1992, Drew 1995, Gayle, Berridge, and Davies 2003, Gayle, Murray, and Connelly 2016b, Gayle, Playford, and Connelly 2020, Playford and Gayle 2016, Playford et al. 2016, Stopforth, Gayle, and Boeren 2021, Strand 2014, Sullivan 2001). In a broader theory of cultural and social reproduction, Bourdieu argued that pupils from more advantaged social class backgrounds have greater stocks of cultural capital. Cultural capital is transmitted between generations within families, and it is converted into more favourable academic outcomes (Bourdieu 1973). In this study, we
investigate the roles that both parental social class and cultural capital have on young people’s outcomes in British school qualifications.

The General Certificates of Secondary Education (GCSEs) are the main qualifications undertaken by pupils in England in school Year 11 (age 15-16). Outcomes in school GCSEs are investigated in this study. Despite the general trend of improvement in school GCSE outcomes in more recent cohorts, there are persistent parental social class inequalities (see Connelly, Murray, and Gayle 2013, Connolly 2006, Demack, Drew, and Grimsley 2000, Gayle, Murray, and Connelly 2016b, Playford and Gayle 2016, Strand 2014, Stopforth, Gayle, and Boeren 2021).

The Bourdieusian position is that differences in cultural capital provide a theoretical explanation of the ‘unequal scholastic achievements of pupils from different social class backgrounds’ (Bourdieu 1986: 47). Following Lamont and Lareau (1988) and Breinholt and Jaeger (2020), we theorise that cultural capital describes the accumulation of a set of skills, knowledge, attitudes, or behaviours which have been sanctioned by the ‘dominant’ classes in society, and which are then converted into more desirable educational outcomes. In empirical work, this is often considered as familial investment in certain ‘middle class’ activities, such as going to the theatre or reading for pleasure, which may aid scholastic achievement. Following Lareau and Weininger (2003), we initially theorise that increased stocks of cultural capital can help pupils to access ‘scarce rewards,’ such as the highest grades in school GCSEs. In this study, through multivariate statistical analyses, we investigate whether measures of cultural capital provide additional explanations of inequalities in school GCSE outcomes. It is theoretically plausible that the occupation-based social class inequalities that were observed in earlier studies may be attenuated when appropriate measures of cultural capital are included in multivariate analyses, and we evaluate this conjecture.

A major challenge when examining the roles of parental social class and cultural capital in children’s school qualifications is locating a large-scale nationally representative data source that includes the necessary measures. The British birth cohort studies 1946, 1958, and 1970 all pre-date the introduction of modern school qualifications (see Wadsworth et al. 2006, Power and Elliott 2006, Elliott and Shepherd 2006). The Longitudinal Study of Young People in England (LSYPE) is an important study but provides only a single cohort of data on young people who were aged 16 in 2006.¹ The Youth Cohort Study of England and Wales (YCS) was a specialist data resource containing detailed information on young people (see Croxford, Ianelli, and Shapira 2007). The YCS was very well suited to studying young peoples’ experiences at school, in training, and the labour market. The YCS contained few candidate measures of cultural capital and did not collect data from parents. The YCS has since been discontinued.² When undertaking these analyses, the Millennium Cohort Study (age 17 sweep) was not yet publicly available to researchers. However, the MCS has the potential to provide data suitable for replicating and extending the analyses in this paper.

Analysing administrative data is a potentially feasible strategy in the absence of suitable social survey data (Jones and Elias 2006). The National Pupil Database (NPD) and the Pupil Level Annual School Census (PLASC) provide important data on school qualifications in England (see Florian et al., 2004). The NPD and PLASC resources do not include measures that are suitable indicators of cultural capital. Indeed, it is common for administrative datasets to contain only a limited set of explanatory variables (Connelly, Gayle, and Lambert 2016b).
To address this issue, we take the innovative step of analysing data extracted from the National Pupil Database (NPD) which is linked to data from Understanding Society, the UK Household Longitudinal Study (UKHLS). The UKHLS is an ongoing large-scale, nationally representative study (Buck and McFall, 2012). The UKHLS collects panel data (i.e. via repeated contacts) on a broad range of subjects, from both adults and young people (Fumagalli, Knies, and Buck 2017).

The cardinal challenge in studies of cultural capital that uses existing survey data is the identification of survey variables that are candidate measures of cultural capital. Sullivan (2002) warns of the difficulties of developing suitable indicators of cultural capital from existing survey datasets. The UKHLS is a general (or omnibus) study and collects an extensive portfolio of measures on the key domains of social and economic life (Platt et al. 2021). Developing a valid and reliable set of measures of cultural capital from data collected in the UKHLS is a sizeable undertaking. We navigate this important challenge by presenting results from two alternative methods of operationalising cultural capital.

**Measuring cultural capital**

A central challenge for this study is developing a valid and reliable set of measures of cultural capital from Understanding Society, the UK Household Longitudinal Study (UKHLS). Several authors have highlighted that Bourdieu’s original definition of cultural capital, and its subsequent measurement, lacked clarity (Lamont and Lareau 1988, Sullivan 2001, Davies and Rizk 2018, Breinholt and Jaeger, 2020). This presents a sizeable, and non-trivial, operationalisation challenge (Kingston 2001, Tan 2017). Sullivan (2002) warns of the problems that follow Bourdieu’s lack of definitional clarity, and further comments that many studies appear to take the convenient course of defining cultural capital in relation to the measures that are most readily available within a research dataset.

In studies of cultural capital, a variety of observed (or manifest) measures have been used to operationalise the concept. Most notably, **parental education level** (Jonsson 1987, Jaeger and Holm 2007, Kraaykamp and Van Eijck 2010); **highbrow cultural participation or cultural knowledge** (Dimaggio 1982, De Graaf 1986, Katsillis and Rubinson 1990, Kalmijn and Kraaykamp 1996, Aschaffenburg and Maas 1997, Roscigno and Ainsworth-Darnell 1999, De Graaf, De Graaf, and Kraaykamp 2000, Sullivan 2001, Jaeger, 2009, Kraaykamp and Van Eijck 2010, Jaeger and Breen 2016); and **reading behaviours** (De Graaf 1986, Kalmijn and Kraaykamp 1996, De Graaf, De Graaf, and Kraaykamp 2000, Sullivan 2001, Dumais 2002, Jaeger and Holm 2007, Jaeger and Breen 2016, Breinholt and Jaeger, 2020).

It is rare for any two empirical studies to use precisely the same measures of cultural capital, and alternative forms of measurement are often used. This gives rise to the inevitable problem that the results are often inconsistent. For example, Dimaggio (1982) operationalised highbrow cultural participation as a factor score of children’s participation in cultural activities like performing arts, visual arts, attending arts events or reading literature. The results demonstrated a strong, positive impact of cultural capital on high school grades in the USA. By contrast, Katsillis and Rubinson (1990) found no association between a factor score of cultural participation (including attending theatres, lectures, museums, and galleries) and high school grades in Greece. Aschaffenburg and Maas (1997) computed a dummy variable of ‘active investment’ in highbrow cultural participation, such as attending classes in music, visual arts, performance, art appreciation, or music appreciation. They
computed separate dummy variables at three stages: before the age of 12, between the age of 12 and 17, and between the age of 18 and 24. The results suggested that cultural capital was strongly associated with school outcomes in the USA, but cultural capital did not have a cumulative effect across the educational life course. Other studies have focused on reading as a sole indicator of cultural capital. For example, Graetz (1988) used a measure of the number of books in the home, which was positively, but mildly, associated with educational transitions net of the effects of social background.

Where studies have operationalised cultural capital as both highbrow participation and reading behaviours, the results have generally found that highbrow participation is not associated with educational outcomes, whereas reading behaviours are. Breinholt and Jaeger (2020) operationalised cultural participation as three summated scales: first, performing arts classes (i.e. dance, music, art, performing arts, and drama classes); second, reading behaviours (i.e. frequency of reading outside of school and number of children’s books); and third, participation in organised clubs or recreational programmes and organised athletics. The authors found that reading behaviours had a direct, positive association with educational performance, whereas other elements of cultural capital did not. De Graaf, De Graaf, and Kraaykamp (2000) derived scales of parental reading behaviours (such as whether parents read regional or historical novels; thrillers, science fiction, or war novels; Dutch literature; translated literature; or literature in a foreign language) in addition to measures of parental beaux arts participation (such as visits to art museums, historical museums, opera or ballet performances, classical music concerts, or theatre performances). The results demonstrated that parental reading behaviours were associated with children’s educational attainment (highest level of education), but cultural participation was not. Sullivan (2001) developed more detailed measures of reading (i.e. the type of books and newspapers read, number of books read, and use of libraries), as well as separate indicators of highbrow cultural participation (i.e. attendance at art galleries, theatres, and concerts). In addition, Sullivan also developed measures of activities such as listening to music and playing an instrument, the types of television programmes watched, a test of cultural knowledge (testing knowledge of famous cultural figures), and language tests (active and passive vocabulary). The results reported that reading behaviours had the greatest effect on GCSE outcomes, and were also associated with cultural knowledge and linguistic tests.

In addition to using different measures, the focus on different analytical units also hinders comparability across studies of cultural capital and education. For example, De Graaf (1986), De Graaf, De Graaf, and Kraaykamp (2000), De Graaf (1988), Jonsson (1987), and Kalmijn and Kraaykamp (1996) focused on parents’ cultural capital. Dimaggio (1982), Graetz (1988), Farkas et al. (1990), Katsillis and Rubinson (1990), Roscigno and Ainsworth-Darnell (1999), Dumais (2002) and Jaeger and Holm (2007) focused on children's cultural capital. Whereas, Aschaffenburg and Maas (1997), Sullivan (2001), Jaeger (2009), Kraaykamp and Van Eijck (2010), Jaeger and Breen (2016) included measures relating to both parents and their children.

Despite the concept of cultural capital being widely deployed as an explanation for educational inequalities, this short review should illustrate that currently there is no consensus on the indicators that are valid and reliable measures of cultural capital. This short review should also highlight that empirical studies have focused on different components of cultural capital. In this study, we explore data from a battery of potential measures of cultural capital that are collected periodically from parents in the UKHLS adult survey and from young people in the UKHLS youth survey.
Data

We analyse data from *Understanding Society*, the UK Household Longitudinal Study (UKHLS) (University of Essex, Institute for Social and Economic Research, NatCen Social Research and Kantar Public 2018). The UKHLS is a major longitudinal study which follows participants over time, regularly collecting data about individuals and their households (Buck and Mcfall 2012). The UKHLS has a very large sample and therefore data collection requires a wide fieldwork window (see Table 1 in Knies 2018). An important innovative feature of the UKHLS is that data from administrative education records have been linked to UKHLS individuals (Department for Education, University of Essex, Institute for Social and Economic Research and NatCen Social Research 2015).

The administrative educational data linked to the UKHLS spans the school academic years 2009/10 to 2012/13. These data are for English state school pupils that are held in the National Pupil Database (NPD). This provides broad coverage as only a small percentage (approximately 7%) of English pupils attend independent schools (Department for Education 2019). A special license is required to access these data, and they must be analysed in a Secure Lab setting. These analyses are subject to Statistical Disclosure Control (SDC) which is the formal process by which the UK Data Service review work to ensure that results cannot be used to identify an individual respondent.

In the UKHLS, there is a battery of potential measures of cultural capital that are collected periodically, as part of a rotating schedule of questions. From wave 2 these questions are asked at every other wave in the youth survey, and every three waves in the adult survey. We analyse the wave 2 responses for both the young person and their parents, and the young person’s subsequent GCSE outcomes from the administrative educational data. The analytical sample comprises 736 young people.

Measures

**School GCSE outcomes**

GCSEs were introduced in the 1980s, and remain the standard qualifications that are undertaken by pupils in England and Wales in school Year 11 (aged 15-16) (Department for Education 1985, Mobley et al. 1986, North 1987). GCSEs have historically been a mixture of assessed coursework and examinations (see Ashford, Gray, and Tranmer 1993). It is usual for pupils to study about nine GCSE subjects (Carroll and Gill 2017). The diet of subjects will include core subjects (e.g. English, Maths and Science) and a subset from a range of non-core subjects (e.g. History and French) (Jin, Muriel, and Sibieta 2011). Generally, each subject is assessed separately and a subject-specific GCSE is awarded. Grades are alphabetical, with the highest grade A* and the lowest grade G (Yang and Woodhouse 2001).

It is common for researchers to score pupils’ school GCSE outcomes (see Croxford, Ianelli, and Shapira 2007, Gayle, Murray, and Connelly 2016b, Yang and Woodhouse 2001, Connolly 2006). Research using data that pre-dated or overlapped the introduction of the A* grade in 1994 tended to combine A* and A grades in their scoring system (Demack, Drew, and Grimsley 2000, Yang and Woodhouse 2001, Gayle, Murray, and Connelly 2016a). For the cohorts of young people in our analyses, the A* grade was well-established and therefore we distinguish between the A* and A grades in the points score. Yang and
Woodhouse (2001) reported that it is standard practice to assign 8 points to each A* grade. We assign an A* grade 8 points, an A grade 7 points, a B grade 6 points, a C grade 5 points, a D grade 4 points, an E grade 3 points, an F grade 2 points, and a G grade 1 point. Unclassified U grades are assigned 0 points. The educational outcome variable analysed in this study is the young person's total school GCSE points.

**Parental social class**

In this study, we measure parental social class using the UK National Socio-Economic Classification (NS-SEC) (Pevalin and Rose 2002). NS-SEC is an occupation-based social class measure (Rose, Pevalin, and O’reilly 2005, Connelly, Gayle, and Lambert 2016). It is associated with three important aspects of individuals’ economic lives; their income security, their short-term income stability, and their longer-term income prospects (Goldthorpe and McKnight 2004). NS-SEC is the official UK social class measure (Office for National Statistics 2010). A great deal of empirical testing has been directed towards developing this classification scheme, and there are agreed and documented standards for using the measure in social research (Office for National Statistics 2010).

**Cultural capital**

There are a variety of questions in the UKHLS which collect measures that can initially be considered as potential measures of cultural capital. There is a well-established distinction between components of cultural capital that indicate highbrow cultural participation, and other components that indicate reading behaviours (for example, see Mikus, Tieben, and Schober 2020, Breinholt and Jaeger, 2020, Sullivan 2001, De Graaf, De Graaf, and Kraaykamp 2000, Crook 1997). We have highlighted that previous studies have used different analytical units. Therefore, in this study we take the sensible precaution of identifying potential indicators of both parental and children's engagement in highbrow culture, and both parental and children's reading activities. The fourteen indicators are displayed in Table 1. The measures are included as dichotomous variables in the analyses that follow.

In these analyses, we measure parental education using a small number of broad categories that generally mirror qualification levels. Following Bukodi and Goldthorpe (2013), we theorise parental education as a measure of socio-economic background which is distinct

| Table 1. Indicators of parents’ and children’s cultural capital. |
|---------------------------------------------------------------|
| Parental Highbrow Cultural Participation | Children’s Highbrow Cultural Participation |
| Parent goes to the theatre | Child goes to the theatre |
| Parent goes to museums/art galleries | Child goes to museums/art galleries |
| Parent visits historical places | Child visits historical places |
| Parent goes to the opera/operetta | |
| Parent goes to classical music performances | |
| **Parental Reading Activities** | **Children’s Reading Activities** |
| Parent reads for pleasure | Child reads for pleasure |
| Parent goes to the library | Child goes to the library |
| Parent is a member of a book club | Child discusses books at home |

In the youth survey, children are asked how many books they have read for pleasure in the past month. We constructed a binary measure of child reading for pleasure (1= read 1 or more books in the past month). In the adult interview, parents are asked if they read for pleasure (not including newspapers, magazines, or comics).
from parental social class. We are aware that a number of other studies have used parental education level as a proxy measure of cultural capital (e.g. Jonsson 1987, Jaeger and Holm 2007, Kraaykamp and Van Eijck 2010). It is a reasonable theoretical assumption that these broad education levels are likely to map onto levels of parental cultural capital. However, the associations between the levels of parental education and the manifest indicators of cultural capital are modest (i.e. values of Cramer’s V range from .01 to .38). We theorise that most British people complete their formal education in early adulthood. This leads us to conjecture that whilst parental education and cultural capital are linked, measures that reflect the activities that parents currently engage in are more important in understanding the effects of cultural capital in the pupil’s family environment. In this analysis, we examine both the effects of parental education and manifest cultural capital measures.

In the presence of multiple indicators of parental and children’s cultural capital, approaches to measurement that foster parsimony in statistically orientated data analyses are desirable. Robson (2003) reminds us that the results of work operationalising the concept of cultural capital are directly contingent upon the way it is defined and measured. We investigate cultural capital using two distinctive methodological approaches. The first approach is the construction of a summated measure of participation in activities related to cultural capital (see Breinholt and Jaeger, 2020). For this approach, we employ a theoretical distinction between highbrow cultural activities and reading behaviours a priori, which is guided by previous empirical work. The second methodological approach is to reduce the dimensionality of the data using factor analysis (see, for example, Dimaggio 1982, Katsillis and Rubinson 1990, Hartas 2016). For this approach, we employ statistical methods to derive the alternative aspects of cultural capital from a set of manifest variables. Both methodological approaches have been used in previous empirical work, but they are seldom reported in the same study. An innovative feature of this study is that we compare and contrast both analytical approaches.

The first data analytical approach develops summated measures of activities related to cultural capital. Following the theoretical distinction between highbrow cultural activities and reading activities, we construct four summated measures (i) parental highbrow participation (ii) children’s highbrow participation (iii) parental reading activities (iv) children’s reading activities. These measures are the total number of highbrow activities (i.e. goes to the theatre, museum, or historic places) or reading activities (i.e. reads for pleasure, goes to the library, or discusses books) undertaken by either parent, or by the child.

The second analytical approach uses factor analysis. This provides a multivariate method for reducing dimensionality, where the information contained in the interrelationships of many variables can be conveyed, to a good approximation, by a much smaller set (Bartholomew, Knott, and Moustaki 2011). Factor analysis is common in analyses of cultural capital, but other latent variable models could also be used (for a review, see Bartholomew 2014). The factor analysis was undertaken using Stata version 15 (Statacorp 2017). The tetrachoric correlation matrix is reported in Table A1 in the appendix.

The analysis produced a three factor solution. Figure 1 depicts the three factors, and reports the rotated factor loadings (see Table A2 in the appendix for the full results). Factors 1 and 3 broadly match the conceptual distinction, prevalent in previous analyses, that distinguishes between measures which communicate culture or status (i.e. highbrow cultural participation), and measures which develop analytical and cognitive skills (i.e. reading behaviours) for parents (Crook 1997, De Graaf, De Graaf, and Kraaykamp 2000, Sullivan
By contrast, the factor analytical model recovers a single factor that represents children’s cultural capital (Factor 2), and does not distinguish between highbrow cultural participation and reading behaviours. In the next stage of the analysis, we derived factor scores using the regression method (see Thomson 1951). The factor scores are standardised (mean = 0; standard deviation = 1).

Results

Table 2 presents the associations between parental social class and the individual, manifest cultural capital measures. There are different levels of parental and children’s participation in highbrow cultural activities across the social class categories (NS-SEC). It is notable that there is not a significant relationship between children’s reading activities and their parental

![Factor analysis of cultural capital measures (rotated factor loadings).](image)

**Table 2.** Associations between parental social class (NS-SEC) and cultural capital measures.

|                          | Cramer’s V | Chi-square (d.f.) |
|--------------------------|------------|-------------------|
| **Parental Highbrow Cultural Participation** |            |                   |
| Parent goes to the theatre | 0.37       | 98.46 (8) ***      |
| Parent goes to museums/galleries | 0.42       | 131.92 (8) ***     |
| Parent visits historical places | 0.42       | 128.11 (8) ***     |
| Parent goes to the opera/operetta | 0.17       | 21.94 (8) **       |
| Parent goes to classical music performances | 0.27       | 52.11 (8) ***      |
| **Children’s Highbrow Cultural Participation** |            |                   |
| Child goes to the theatre | 0.22       | 35.38 (8) ***      |
| Child goes to museums/galleries | 0.23       | 37.57 (8) ***      |
| Child visits historical places | 0.24       | 42.42 (8) ***      |
| **Parental Reading Activities** |            |                   |
| Parent reads for pleasure | 0.30       | 66.93 (8) ***      |
| Parent goes to the library | 0.17       | 20.44 (8) **       |
| Parent is a member of a book club | 0.16       | 18.88 (8) *       |
| **Children's Reading Activities** |            |                   |
| Child reads for pleasure | 0.11       | 8.84 (8)           |
| Child goes to the library | 0.13       | 11.43 (8)          |
| Child discusses books at home | 0.09       | 5.97 (8)           |

Associations calculated on data unweighted and unadjusted for survey design. * p < .05, ** p < .01, *** p < .001.
The respective associations between parental occupational social class (NS-SEC) and parental highbrow activities, children's highbrow activities, and parental reading activities are moderate in strength (i.e. values of Cramer's V from .16 to .42). Associations between parental NS-SEC and cultural capital activities are generally higher for parental measures than for children's measures.

Table 3 reports the summary statistics for pupils' school GCSE scores by parental social class (NS-SEC) and the manifest cultural capital indicators. There are sizable occupation-based social class inequalities. For example, the mean GCSE score for pupils from families in NS-SEC 1.2 (Higher professional occupations) is 28 points higher than the mean for pupils from families in NS-SEC 7 (Routine occupations). Pupils that participate in highbrow cultural activities, or have parents who participate in highbrow cultural activities, have higher mean school GCSE scores. Similarly, pupils who engage in reading activities, or have parents who engage in reading activities, also have higher mean school GCSE scores.

In the first stage of the multivariate analysis, we estimate a standard linear regression model on school GCSE scores. We then evaluate a series of nested regression models (see Allen 1997). In addition to assessing the variance explained, we also consider model parsimony using a Bayesian information criterion statistic (BIC) (see Raftery 1986, Treiman 2009). The UKHLS has a complex structure and sample (Lynn 2009). The complexity of the UKHLS data has been appropriately incorporated into the analyses using Stata's survey data function (Kreuter and Valliant 2007). The regression model results (e.g. coefficients and standard errors) are derived from models that are adjusted for the complex survey design and selection strategy in the UKHLS. It is not currently possible to derive easily comparable goodness of fit and parsimony measures for regression models which have been adjusted for complex survey designs. Therefore, the goodness of fit and parsimony measures that we present compare unadjusted models.

Occupation-based parental social class (NS-SEC) explains 16% of the variance in school GCSE scores (see Model 1, Table 4). The explanation is improved by the inclusion of additional explanatory variables; housing tenure, gender, ethnicity, age, and academic year (Adjusted $R^2 = .22$; see Model 2, Table 4). Including a measure of parental education provides a further parsimonious improvement (Adjusted $R^2 = .27$; see Model 3, Table 4). Therefore, parental education is an important additional socio-economic indicator in these analyses. We are aware that a number of other studies have used parental education level as a proxy measure of cultural capital. Comparing Model 2 and Model 3 in Table 5, it is notable that the inclusion of parental education does not attenuate the strong parental social class effects on school GCSE scores.

The four summated cultural capital measures (parental highbrow participation; children's highbrow participation; parental reading activities; children's reading activities) each provide a marginal increase in the proportion of variance explained (see Models 4-7, Table 4). However, parental highbrow participation and children's highbrow participation are not statistically significant when parental reading activities and children's reading activities are included in the model. The model of best fit (Model 11, Table 4) notably includes occupation-based parental social class, parental reading activities, and children's reading activities (Adjusted $R^2 = .31$ and BIC = 6542.46 at 28 degrees of freedom).

The three cultural capital factors (Factor 1 parental highbrow factor score; Factor 2 children's factor score; Factor 3 parental reading factor score) each provide a marginal increase in the proportion of variance explained (see Models 8-10, Table 4).
Table 3. Summary statistics for parental social class (NS-SEC), cultural capital measures and school GCSE scores.

| Parental Social Class | Freq. | Percent | Mean GCSE score | Std. Dev. |
|-----------------------|-------|---------|-----------------|-----------|
| **Parental NS-SEC**   |       |         |                 |           |
| 1.1 Large employers and higher managerial | 42    | 5.71    | 53.31           | 18.35     |
| 1.2 Higher professional occupations | 64    | 8.70    | 62.08           | 16.11     |
| 2 Lower management/professional occupations | 151   | 20.52   | 49.40           | 19.30     |
| 3 Intermediate occupations | 76    | 10.33   | 45.25           | 20.41     |
| 4 Small employers and own account workers | 73    | 9.92    | 44.23           | 21.43     |
| 5 Lower supervisory and technical occupations | 43    | 5.84    | 39.26           | 19.89     |
| 6 Semi-routine occupations | 88    | 11.96   | 38.65           | 23.60     |
| 7 Routine occupations | 56    | 7.61    | 33.75           | 21.84     |
| 8 Not in employment | 143   | 19.43   | 31.55           | 21.01     |

| Parental Highbrow Cultural Participation |       |         |                 |           |
| Parent goes to the theatre |       |         |                 |           |
| Yes | 352 | 47.83 | 49.36 | 21.16 |
| No | 384 | 52.17 | 37.64 | 21.81 |

| Parent goes to museums/ galleries |       |         |                 |           |
| Yes | 336 | 45.65 | 49.38 | 21.07 |
| No | 400 | 54.35 | 38.09 | 21.97 |

| Parent visits historical places |       |         |                 |           |
| Yes | 485 | 65.90 | 47.93 | 21.35 |
| No | 251 | 34.10 | 34.20 | 21.23 |

| Parent goes to the opera/operetta |       |         |                 |           |
| Yes | 27  | 3.67  | 53.82 | 17.44 |
| No | 709 | 96.33 | 42.85 | 22.35 |

| Parent goes to classical music performances |       |         |                 |           |
| Yes | 79  | 10.73 | 59.76 | 16.88 |
| No | 657 | 89.27 | 41.26 | 22.03 |

| Children's Highbrow Cultural Participation |       |         |                 |           |
| Child goes to the theatre |       |         |                 |           |
| Yes | 380 | 51.63 | 48.46 | 21.33 |
| No | 356 | 48.37 | 37.68 | 21.93 |

| Child goes to museums/ galleries |       |         |                 |           |
| Yes | 386 | 52.45 | 49.24 | 22.74 |
| No | 350 | 47.55 | 36.64 | 19.76 |

| Child visits historical places |       |         |                 |           |
| Yes | 366 | 49.73 | 48.73 | 22.52 |
| No | 370 | 50.27 | 37.82 | 20.67 |

| Parental Reading Activities |       |         |                 |           |
| Parent reads for pleasure |       |         |                 |           |
| Yes | 579 | 78.67 | 45.98 | 21.41 |
| No | 157 | 21.33 | 33.16 | 22.55 |

| Parent goes to the library |       |         |                 |           |
| Yes | 371 | 50.41 | 48.39 | 21.49 |
| No | 365 | 49.59 | 38.02 | 21.86 |

| Parent is a member of a book club |       |         |                 |           |
| Yes | 33  | 4.48  | 55.67 | 21.95 |
| No | 703 | 95.52 | 42.66 | 22.13 |

| Children's Reading Activities |       |         |                 |           |
| Child reads for pleasure |       |         |                 |           |
| Yes | 528 | 71.74 | 46.53 | 22.14 |
| No | 208 | 28.26 | 34.91 | 20.40 |

| Child goes to the library |       |         |                 |           |
| Yes | 335 | 45.52 | 48.33 | 22.10 |
| No | 401 | 54.48 | 39.01 | 21.54 |

| Child discusses books at home |       |         |                 |           |
| Yes | 205 | 27.85 | –     | –    |
| No | 531 | 72.14 | –     | –    |

**Total n** 736 100 43.25 22.27

Summary statistics are unweighted and unadjusted for survey design. Mean GCSE score and standard deviation not available for child discusses books at home due to the statistical disclosure control protocol.
The highbrow factor score (Factor 1) is not statistically significant when children's factor score (Factor 2) and parental reading factor score (Factor 3) are also included in the model. The model of best fit (Model 12, Table 4) notably includes occupation-based parental social class, children's factor score (Factor 2), and parental reading factor score (Factor 3) (Adjusted R² .32 and BIC = 6512.89 at 24 degrees of freedom).

The linear regression results for the four substantively important models (Models 2, 3, 11, and 12) are reported in Table 5. An examination of the coefficients and standard errors for parental NS-SEC indicate that the cultural capital measures do not attenuate the substantial parental social class inequalities in school GCSE outcomes. We note that the coefficients for the categories of parental social class (NS-SEC) are attenuated when parental education is also included in the model, however growing up in a family from a less
Table 5. Linear regression models of school GCSE outcomes.

| Linear Regression: GCSE Point Score | Model 2 | Model 3 | Model 11 | Model 12 |
|-----------------------------------|---------|---------|----------|----------|
| **Parental NS-SEC**               |         |         |          |          |
| 1. Large employers and higher managerial occupations | −9.74 (4.71) * | −9.01 (4.70) * | −8.09 (4.66) * | −8.50 (4.78) * |
| 1.2 Higher professional occupations | 0.00 | 0.00 | 0.00 | 0.00 |
| **2 Lower management and professional** | −13.10 (3.24) *** | −9.55 (3.24) ** | −9.61 (3.19) ** | −9.41 (3.19) ** |
| **3 Intermediate occupations** | −16.53 (3.42) *** | −9.51 (3.56) ** | −9.42 (3.66) * | −8.80 (3.73) * |
| **4 Small employers and own account workers** | −16.72 (4.10) *** | −10.20 (4.12) * | −10.05 (4.26) * | −9.28 (4.16) * |
| **5 Lower supervisory and technical** | −21.61 (4.37) *** | −14.01 (4.32) ** | −12.02 (4.14) ** | −11.38 (4.13) ** |
| **6 Semi-routine occupations** | −21.71 (4.23) *** | −12.94 (4.53) ** | −12.76 (4.46) ** | −12.74 (4.41) ** |
| **7 Routine occupations** | −24.41 (4.19) *** | −14.74 (4.40) ** | −15.39 (4.37) *** | −14.04 (4.35) ** |
| **8 Not in employment** | −25.22 (3.80) *** | −16.87 (4.08) ** | −15.75 (4.03) ** | −15.27 (4.00) *** |
| **Housing Tenure**               |         |         |          |          |
| Owned/private rented             | 0.00 | 0.00 | 0.00 | 0.00 |
| Social housing                   | −13.57 (2.52) *** | −11.70 (2.54) *** | −10.44 (2.36) *** | −10.07 (2.36) *** |
| **Gender**                       |         |         |          |          |
| Male                              | 0.00 | 0.00 | 0.00 | 0.00 |
| Female                           | 7.40 (1.90) *** | 7.05 (1.82) *** | 4.75 (1.85) * | 5.03 (1.84) ** |
| **Ethnicity**                    |         |         |          |          |
| White                             | 0.00 | 0.00 | 0.00 | 0.00 |
| Mixed                            | 1.82 (3.83) | −0.93 (4.27) | −0.66 (4.26) | 0.63 (4.51) |
| Asian/Asian British              | 2.78 (4.45) | 0.80 (4.71) | 0.72 (4.21) | 1.29 (4.18) |
| Black/Black British              | 3.85 (4.49) | 2.77 (4.45) | 1.72 (4.05) | 2.71 (4.16) |
| Other                            | 12.83 (6.81) | 9.80 (7.25) | 13.29 (7.94) | 11.88 (7.17) |
| **Parental Education Level**     |         |         |          |          |
| Higher education                 | 0.00 | 0.00 | 0.00 | 0.00 |
| Further education                | −7.86 (2.94) ** | −6.18 (2.90) * | −6.82 (2.80) * | |
| School-level education           | −12.88 (2.42) *** | −10.88 (2.52) *** | −10.87 (2.32) *** | |
| Below school-level education     | −16.21 (3.51) *** | −12.24 (3.70) ** | −12.17 (3.59) ** | |
| **Parental Reading Activities**  |         |         |          |          |
| No activities                    | 0.00 | 0.00 | 0.00 | 0.00 |
| 1 activity                       | 7.31 (3.53) | 7.31 (3.53) | 7.31 (3.53) | 7.31 (3.53) |
| 2 activities                     | 10.71 (3.48) ** | 10.71 (3.48) ** | 10.71 (3.48) ** | 10.71 (3.48) ** |
| 3 activities                     | 12.67 (5.17) * | 12.67 (5.17) * | 12.67 (5.17) * | 12.67 (5.17) * |
| **Children's Reading Activities** |         |         |          |          |
| No activities                    | 0.00 | 0.00 | 0.00 | 0.00 |
| 1 activity                       | 3.92 (2.22) | 3.92 (2.22) | 3.92 (2.22) | 3.92 (2.22) |
| 2 activities                     | 7.68 (2.56) ** | 7.68 (2.56) ** | 7.68 (2.56) ** | 7.68 (2.56) ** |
| 3 activities                     | 9.00 (2.93) ** | 9.00 (2.93) ** | 9.00 (2.93) ** | 9.00 (2.93) ** |
| **Children's Cultural Capital Factor Score** | 3.92 (1.05) *** | 3.92 (1.05) *** | 3.92 (1.05) *** | 3.92 (1.05) *** |
| **Parental Reading Factor Score** | 2.70 (1.16) * | 2.70 (1.16) * | 2.70 (1.16) * | 2.70 (1.16) * |
| **Observations**                 | 736 | 736 | 736 | 736 |
| **Goodness of fit measures**     |         |         |          |          |
| Adjusted R²                      | 0.22 | 0.27 | 0.31 | 0.32 |
| BIC (df.)                        | 6579.38 (19) | 6550.59 (22) | 6542.46 (28) | 6512.80 (24) |
| BIC (based on deviance)          | 1720.87 | 1692.08 | 1683.96 | 1654.30 |
| BIC (based on likelihood ratio chi square) | −82.93 | −111.72 | −119.84 | −149.50 |

Source: Wave 2 UKHLS youth questionnaires with linked NPD records. Adjusted for complex survey design and selection strategy. Models also include age and academic year. *p < .05, **p < .01, ***p < .001. The model constants are omitted due to statistical disclosure control protocols.
advantaged social class position still has a large negative effect. There is no evidence that the occupation-based parental social class effects that were observed in earlier studies were only significant because occupational social class acted as a proxy in the absence of appropriate measures of cultural capital.

Occupation-based parental social class is an important element of these inequalities in school GCSE outcomes. The measures of cultural capital provide a small additional explanation of inequalities in school GCSE outcomes. Cultural capital is unequally distributed across the occupation-based social class categories, but the associations are relatively weak. Parental and children’s reading activities are significant, but they only explain an additional 4-5% of the variance in school GCSE outcomes. In a multivariate context, reading activities play a much less influential role than occupation-based social class in explaining social inequalities in school GCSE outcomes. Higher quantities of parental reading activity are beneficial, both when measured by a direct measure of reading activities or a latent factor. Children’s reading activities are also important. Engaging in two or three reading activities, on average, increases the pupil’s GCSE score by between seven and nine points. The size of this effect should not be overlooked since an extra GCSE pass at grade A* is worth eight points. Taken together, these results do not lend any support to the theoretical position that parental occupation-based social class inequalities in school GCSE outcomes are attenuated when detailed measures of cultural capital are included in the analyses.

Conclusions

The persistence of social class inequalities in school GCSE outcomes is disturbing because qualifications gained at school are consequential, and affect individual young people’s futures in both education and employment. Using multivariate statistical analyses, we investigated the role of both occupation-based social class and cultural capital, and conclude that measures of cultural capital provide little additional explanation of inequalities in school GCSE outcomes. A priori, it was theoretically plausible that the occupation-based social class inequalities that were observed in earlier studies may be attenuated when appropriate measures of cultural capital are included in multivariate analyses. This conjecture is not supported in these empirical analyses. The family’s location in the occupational class structure remains salient. As Bihagen and Lambert (2018) remind us, the occupations that are located in more advantaged social classes tend to be more highly skilled and tend to have more highly educated incumbents.

The empirical analyses convey the important message that engagement in highbrow cultural activities is not important, but reading activities are influential. This finding chimes with other results on the effects of cultural capital on educational outcomes (Crook 1997, De Graaf, De Graaf, and Kraaykamp 2000, Mikus, Tieben, and Schober 2020, Sullivan 2001). Breinholt and Jaeger (2020) reasoned that highbrow cultural participation might signal scholastic ability to teachers, but reading behaviours develop relevant educational skills. Sullivan (2001, 2008) similarly noted that there is an important difference between language based activities (e.g. reading) and non-language based activities (e.g. highbrow cultural participation). Sullivan and Brown (2015) argued that the ability to process language is a key skill in education, which has direct applications to scholastic work. The influence of parental reading behaviours may operate through ‘passive role modelling’ (Mikus, Tieben, and Schober 2020). It is plausible that children living in households in
which adults actively engage in reading activities may be more likely to engage in reading activities themselves (De Graaf, De Graaf, and Kraaykamp 2000).

Many education and youth datasets include measures of parental education but do not have suitable measures of cultural capital. We consider that parental education level is a useful socio-economic indicator that reflects the educational milieu of the family, but we strongly argue it is not a suitable substitute for cultural capital. This is because a higher level of parental education does not automatically engender a greater stock of cultural capital. Cultural capital measures should reflect what parents and children do, not just the qualifications parents gained in early adulthood. In this study, the persistent significance of parental education, even when detailed measures of cultural capital are included in the models, persuades us that whilst parental education is associated with cultural capital, it should not be used interchangeably. Therefore, we warn against the uncritical use of parental education as a measure of cultural capital.

This study has directly addressed the challenges associated with operationalising measures of cultural capital from an existing data resource. On reflection, the use of survey data with linked educational administrative data has proved beneficial, and addressed the data challenges associated with studying parental social class, cultural capital, and educational outcomes. We are cognisant that the analytical sample is smaller than samples that are common in the older UK birth cohort studies and the discontinued Youth Cohort Study of England and Wales. Since undertaking these analyses, a new secure access dataset has been released linking English education data (including GCSE exam results) to the participants of the Millennium Cohort Study (Centre for Longitudinal Studies 2019). These data are likely to prove valuable for studying school qualifications for a larger cohort of school pupils. This current study provides a useful methodological blueprint which will inform our future analyses of the Millennium Cohort Study.

The results reported above make an important contribution to understanding the role of cultural capital in educational outcomes. Increasing levels of cultural capital is a positive valance concept, and it is beguiling to believe that increasing pupils’ levels of cultural capital will have a positive influence on school GCSE outcomes. It is tempting to theorise that visits to museums or historic venues might be helpful in igniting interests in history, and that visits to the theatre might similarly cultivate learning in drama. On deeper reflection, it is difficult to plausibly describe mechanisms by which the exposure to certain extracurricular activities would influence outcomes in other school GCSE subjects. Indeed, to our knowledge, educational commentators seldom (if ever) suggest that going to football matches or attending church has any positive effects on outcomes in GCSE Physical Education or GCSE Religious Studies. The equivocal theoretical claims about the mechanisms by which cultural capital might influence school GCSE outcomes, is contrasted by the far more conceivable theoretical view that young people engaging in reading behaviours is a mechanism that can develop educational skills that are pertinent to the wider portfolio of school GCSE outcomes.

The concept of cultural capital has made the transition from social science to educational policy. New guidance from the Office for Standards in Education, Children’s Services and Skills (Ofsted) states that ‘as part of making the judgement about quality of education, inspectors will consider the extent to which schools are equipping pupils with the knowledge and cultural capital they need to succeed in life’ (Office for Standards in Education Children’s Services and Skills, 2019: 10). Leading British sociologist John Goldthorpe
commented that the ‘the whole passage [in the handbook] is so vacuous as to be quite innocuous’9 The definitions in the policy are vague, and we are convinced that the assessment of cultural capital will be plagued by measurement problems.

This study reports a set of empirical findings that do not lend any support to the view that increasing cultural capital will reduce the size of social class inequalities in school GCSE outcomes. This is not to argue that activities that have sometimes been associated with increasing cultural capital should not be part of the school experience, for example extracurricular trips may contribute to educational enjoyment. However, if schools are serious about reducing educational inequalities, our empirical findings send a clear and actionable message for policy and practice. In order to support groups of pupils who may be projected to have lower school GCSE outcomes, schools would be better placed to concentrate on increasing reading activities. Our findings further demonstrate that this would be in addition to more directly targeted policy interventions to reduce the stark social class attainment gaps, which are consistently observed in school GCSE outcomes.

Notes

1. See https://cls.ucl.ac.uk/cls-studies/next-steps/ accessed 6th October 2021.
2. See https://data.gov.uk/dataset/2521ad1-8f52-4fd5-a0f8-3bb0a0453acb/youth-cohort-study accessed 6th October 2021.
3. See https://ukdataservice.ac.uk/find-data/access-conditions/ accessed 6th October 2021.
4. See https://ukdataservice.ac.uk/help/secure-lab/securelab-faqs/#SDC accessed 6th October 2021.
5. Other qualifications are available but the majority of pupils study GCSEs (Gill, 2016).
6. Following reforms from 2016, a new GCSE grading system is in the process of being introduced in England. Grade 9 is the new highest grade and grade 1 the lowest (Ofqual, 2018). Wales has retained the A*-G grading system (see https://www.qualificationswales.org/english/qualifications/gcses-and-a-levels/gcses accessed 6th October 2021). In Northern Ireland, the grading system used varies by exam board (see https://www.nidirect.gov.uk/articles/gcses accessed 6th October 2021).
7. We present the results from a principal components analysis, but we also re-estimated the analysis using the principal factor method and the results are consistent.
8. Parent library use loaded weakly on all factors and child theatre attendance loaded similarly on factor 2 (0.44) and factor 3 (0.45). We re-estimated the analyses separately without these measures and the results were unchanged.
9. See https://www.theguardian.com/education/2019/sep/03/ofsted-plan-inspect-cultural-capital-schools-attacked-as-elitist accessed 6th October 2021.

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