THE ATTITUDINAL LEGACY OF COMMUNIST LABOR RELATIONS

DAVID G. BLANCHFLOWER and RICHARD B. FREEMAN*

This study of workers’ attitudes compares data from International Social Survey Programme (ISSP) surveys for former communist countries in Europe with ISSP data for Western countries over the period 1987–93, which covers the beginning of the transition to a market economy for the former communist countries. Consistent with their hypothesis that communist-run economies left an attitudinal “legacy,” the authors find that the citizens of former communist countries evinced a greater desire for egalitarianism, less satisfaction with their jobs, and more support for strong trade unions and state intervention in the job market and economy than did Westerners. Over the course of the period studied, however, residents of the former communist European countries perceived sizable increases in occupational earnings differentials, and they adjusted their views of the differentials that “ought to” exist in their economies in the direction of greater inequality.

Labor relations in communist economies has diverged historically from that in free market economies. Under communism nearly all workers joined official “transmission belt” unions that operated as an arm of the state rather than as independent representatives of workers. The state set wages, prices, and enterprise budgets in ways that created huge job vacancies with no open unemployment, and produced low real wages and narrow skill and sectoral pay differentials. These practices resulted in inefficient allocations of labor (Freeman 1992). At the same time, there was a social ethos favoring egalitarianism (Burawoy 1985; Kornai 1980). While most analysts believe that communist labor practices produced demoralized and disgruntled workers, there have been no comparisons of worker attitudes in historically communist economies and market economies using comparable survey instruments that document or test this expectation.1

*David Blanchflower is Professor of Economics at Dartmouth College and Research Associate at the National Bureau of Economic Research. Richard Freeman is Professor of Economics at Harvard University, Director of Labor Studies at the National Bureau of Economic Research, and Director of the Programs in Discontinuous Economics at the London School of Economics. The authors thank Alan Krueger for helpful suggestions on an earlier draft of this paper.

1Three papers have used attitudinal data to examine the development of markets in general, but not specifically labor markets, in eastern Europe. Akerlof and Yellen (1991) looked at various attitudes of workers in East Germany and comparable workers in West

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To what extent do workers in traditionally communist societies differ from workers in the West in their attitudes toward working conditions, wage inequality, job satisfaction, and the role of unions and the state in determining labor market outcomes? To what extent can any observed differences be attributed to the "legacy" of the communist past? How have attitudes changed during the transition to a market economy?

In this study we use data from the International Social Survey Programme (ISSP) to try to answer those questions. Each year the ISSP focuses on a particular topic. The topics most relevant to our area of inquiry are "social inequality" (the 1992 and 1987 modules), "work orientation" (the 1989 module, supplemented by preliminary evidence from the 1993 module), and "the role of government" (1990). We have information on seven ex-communist countries: Bulgaria, Czechoslovakia, East Germany, Hungary, Poland, Slovenia, and Russia. Most of these countries reported in the 1992 module. For Hungary and Poland, we have information in several ISSP modules, which allows us to analyze changes in labor outcomes and attitudes over time. Appendix Table A1 reports the number of responses by country and year.

Measuring The Communist Legacy in the Labor Market

Like any other lengthy historical experience, decades of communist rule in eastern Europe and Russia have arguably left a "legacy" of effects on workers and the job market that successor market economies must deal with. There is a potential labor market outcome legacy: narrow wage distributions; distinct work practices (Haraszti 1978; Burawoy and Lukacs 1992); relations with management that diverge from those in historically capitalist economies (Burawoy 1985; Kornai 1980); high rates of union membership in unions dominated by communist regimes; dependence of firms on the state; and enterprise provision of many social benefits. In addition, there is a potential "attitudinal legacy" in terms of workers’ views about institutions and markets due to experience under communism and decades of communist propaganda.

To be sure, institutions, labor market outcomes, and attitudes are likely to differ among ex-communist countries, just as they do among Western countries. Poland and Hungary differ in the nature of their former communist systems and in the characteristics of their movement to a market economy, just as Sweden and Norway differ from the United States and the United Kingdom. National differences notwithstanding, however, the operative hypothesis of this study is that the communist experience was sufficiently similar to leave an identifiable common legacy affecting outcomes and views of the labor market in the ex-communist states. This is a hypothesis about cross-national differences due to comparable experiences, and as such is subject to the problems of assessment common to such analyses (Kohn 1989).

A priori, what form might this legacy take? By imposing narrow wage distributions through governmental fiat, and pressing
an egalitarian ethos, communism should have influenced attitudes toward earnings inequality and government interventions in labor market outcomes. We anticipate three differences in attitudes between workers in ex-communist and Western countries: (1) employees in ex-communist countries will be more favorable toward narrow wage distributions and government interventions to attain them; (2) given the persistence of poor working conditions in former communist countries (in which workers were deprived of normal market modes of responding to such conditions, having neither the "exit" option of finding employment outside the state-run sector nor the "voice" option of forming free trade unions), employees in those countries will evince less job satisfaction than workers in Western countries; and (3) workers in ex-communist countries will have more positive attitudes toward free trade unions than do workers in the West. In addition, assuming that attitudes toward inequality respond to actual levels of inequality, we expect that the views in ex-communist countries toward wage differences will change over time, as competitive markets produce greater inequality.

We examine the ISSP for evidence of such "legacies" in two ways. First, we contrast responses to ISSP questions (both outcome and attitude items) in ex-communist countries with those in Western countries, controlling for respondents' individual characteristics. This is effectively an exercise in estimating country dummy variables: the test of the legacy hypothesis is whether there is a common pattern among the dummy variables for ex-communist countries. Second, we probe the correlates and investigate possible objective causes for any observed differences. This is an effort to explain differences between ex-communist countries and Western countries in terms of actual or perceived labor market factors.

In attitude surveys, the validity and reliability of measures are invariably of some concern. Because none of our data are longitudinal, we cannot determine whether the same person would respond similarly to the same question in a second interview, which would give us test-retest measurement errors. This does not mean, however, that we have no way to assess the quality of responses. We perform three checks on individual responses: (1) we examine whether individual responses fit a pattern—for instance, whether workers who are completely satisfied with their jobs also express more willingness than others to work without pay; (2) we relate individual responses to objective characteristics: are workers who want a more narrow wage distribution high or low in that distribution?; and (3) we investigate whether people with the same characteristics give comparable responses in different surveys. The results of all of these checks suggest that responses on the ISSP are consistent.  

Characteristics of ex-Communist and Western Data Samples

For our cross-section analyses of workers in ex-communist economies and market economies, we rely heavily on the 1992 ISSP, which provides data on seven ex-communist countries (Bulgaria, the Czech Republic, East Germany, Hungary, Poland, Russia, and Slovenia) supplemented by the 1993 ISSP. For our analysis of changes over time, we use the 1987 and 1992 ISSP surveys, which contain comparable information for Hungary and Poland. None of these data files are ideal. Relatively few  

As examples of our checks, we found that 24% of those completely satisfied with their job agreed strongly that they would work without pay, compared to less than 10% of those who were dissatisfied, 13% who were neither satisfied nor dissatisfied, and so on. Also, we confirmed that greater inequality was favored by the more educated (see Table 2) and those with high earnings (table available from authors on request); we found, from a comparison of responses on similar questions in 1987 and 1992, that years of schooling had a comparable effect in both years (see coefficients reported in Table 2); and we confirmed that age, sex, and marital status also had comparable effects in both years. In addition, we explored individual differences in responses to the Table 7 questions and found that people with similar demographic characteristics responded in the same direction.
workers in the ISSP for the ex-communist countries report being in the private sector, which prevents us from making inferences about differences between those employees, who will constitute the bulk of future employment, and employees of state-owned firms. For our comparison group we have data from eleven OECD countries in 1987 or 1992 (Australia, Austria, Canada, Great Britain, Italy, Netherlands, New Zealand, Norway, Switzerland, West Germany, and the United States) plus information on other variables of interest on six other countries (Israel, Japan, the Philippines, Southern Ireland, Spain, and Sweden) in other years of the ISSP.6

We compare the former communist states with a set of Western countries rather than with a single one because differences between the eastern European countries and any particular Western country, such as the United States, could be due to the distinct characteristics of that Western country rather than to the legacy of communist labor practices vis-à-vis market economies in general. The danger of this approach, however, is that we may miss important contrasts between more narrowly defined groups, say between East Germany and West Germany. To deal with this danger, we present all calculations with individual country dummies before giving broader ex-communist versus traditional market economy contrasts.

Attitudes Toward Inequality

As a starting point for analyzing differences in views toward inequality between persons in ex-communist economies and those in traditional capitalist economies, we examined responses to the 1992 ISSP question: “Are differences in income in this country too large?” The difference between the ex-communist countries and the West is stark.7 Save for Italy, there is only the barest overlap in the survey responses. On average, 52% of respondents in the ex-communist countries expressed a strong belief that income differences were too large, compared to 30% of respondents in Western countries. Most striking, perhaps, are the differences between citizens in the ex-communist countries and those in Scandinavia, where incomes are more equally distributed than in most capitalist economies and where the social democratic egalitarian ethos is strong. In Sweden and Norway, just 23–24% of respondents believed income differences were too large.

We pursue the analysis of differences in views toward inequality across countries using two questions on the 1987 and 1992 ISSPs on pay by occupation. In those years the ISSP asked individuals in different countries “what people earn each year” and what they “ought to earn” in 11 occupations. In each of the two years, seven of these occupational groups were identical, while four were different.8 We used the reported figures on what people earn to estimate a measure of “perceived differentials” and used the reported figures on what they ought to earn to estimate “appropriate differentials.”9 We calculated from these re-

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6 In a number of cases the data were collected in a different year from that reported. For example, Great Britain did not conduct a survey in 1988; the reported data were collected in 1989. Half the 1989 respondents were asked the questions from the 1988 ISSP and the other half were asked the questions from the 1989 component. In a few cases, such as Italy in 1988, 1989, 1990, and 1991 and Austria in 1987 and 1988, the same group of respondents were asked the two years of questions.

7 The percentages agreeing strongly that differences in income were too large were as follows: Bulgaria, 85%; East Germany 61%; Russia, 59%; Slovenia, 48%; Poland, 42%; Czech Republic, 36%; Hungary 35%; Italy, 53%; Great Britain, 37%; Austria, 35%; West Germany, 31%; New Zealand, 30%; United States, 28%; Canada, 26%; Sweden, 24%; Norway, 23%; and Australia, 18%.

8 The occupational categories available in both years were doctor, small shopkeeper, chairman of large national company, skilled factory worker, farm worker, unskilled factory worker, and cabinet minister. In 1987 the remaining four categories were bricklayer, bank clerk, secretary, and bus driver, whereas in 1992 they were replaced by solicitor, shop assistant in a department store, owner/manager of a factory, and appeal court judge.

9 Summaries of the mean earnings by country in the two categories are reported in a data appendix that is available on request from the authors.
sponges the standard deviation of the natural log of the averages reported in these appendix tables for each country.

The resultant statistic does not measure what survey respondents believe occupational differentials should be. They were not asked about occupational differentials per se, but rather about occupational pay. But implicit in statements about pay for different occupations are statements about differentials in pay. For instance, someone who reports that the chairman of a large national company earns or should earn X while a bricklayer earns or should earn Y is telling us about the differentials he perceives to exist or believes should exist. The legacy hypothesis suggests that persons in ex-communist countries should differ from those in capitalist countries both in the occupational differentials they perceive to exist and in the differentials they believe to be fair (that is, the differentials they believe “ought to” exist).

Table 1 gives the standard deviations of the natural log of earnings from the 1987 and 1992 ISSP modules. In 1987, workers in the two ex-communist countries for which we have data, Hungary and Poland, perceived smaller occupational differentials than did workers in most Western countries. The unweighted country mean for Hungary and Poland in the standard deviation of perceived differentials is, for example, 0.57, which falls considerably short of the 0.81 for workers in the West and the 0.85 for those in the four large ex-communist countries for which we have data in both 1987 and 1992 (the United States, West Germany, Great Britain, and Italy). The second column in the table shows a similar ex-communist/Western difference in the pattern of occupational pay differentials that people believe ought to exist: persons in the ex-communist states favor smaller differentials than persons in the West. In the 1987 data, the average standard deviation of the natural log of earnings perceived to be fair is 0.40 for Hungary and Poland, compared to an average of 0.60 in the capitalist countries.

By contrast, in 1992, when the transition to a market economy was under way, persons in Hungary and Poland and in the other ex-communist countries for which we have data reported very different perceptions of pay differentials. The differentials they perceived were wider than those perceived by respondents in the Western countries in the ISSP (which includes Norway, with an extremely narrow wage distribution) and similar to those perceived by workers in the big ex-communist countries. Similarly, what Hungarians and Poles believed “ought to” be paid to workers in different occupations was much greater in 1992 than in 1987. Because the “ought to” differentials are also high for workers in the other ex-communist countries, the unweighted mean of the standard deviation of the natural log of earnings perceived to be fair is in the same range for employees in the ex-communist countries as for those in the West: it is slightly higher than the unweighted mean for all Western countries (inclusive of Norway) and slightly lower than the mean for the larger ex-communist countries.

In both years, and in all countries, ex-communist and capitalist, there seems to be a relation between the distribution of perceived pay and the distribution of pay perceived to be fair. The “ought to” differentials are invariably smaller than the perceived differentials, suggesting that workers may form their views of the appropriate pay by occupation on the basis of perceived pay, scaling the latter down by some amount.
Table 1. Perceptions of Actual Earnings and Fair Earnings Across 11 Occupations: Responses of Those Surveyed in Six Ex-Communist and Eleven Capitalist Countries, 1987 and 1992. (Mean Standard Deviation of Log Earnings)

| Country               | 1987 Perceived Earnings | 1987 Earnings Viewed as Fair | 1992 Perceived Earnings | 1992 Earnings Viewed as Fair |
|-----------------------|-------------------------|-------------------------------|-------------------------|-------------------------------|
| Ex-Communist          |                         |                               |                         |                               |
| Bulgaria              |                         |                               | .77                     | .53                           |
| Czechoslovakia        |                         |                               | 1.35                    | 1.07                          |
| East Germany          |                         |                               | 1.03                    | .78                           |
| Hungary               | .59                     | .41                           | .94                     | .67                           |
| Poland                | .54                     | .38                           | 1.02                    | .68                           |
| Russia                |                         |                               | 1.09                    | .71                           |
| **Unweighted Mean**   | **.57**                 | **.40**                       | **1.03**                | **.74**                       |
| Capitalist            |                         |                               |                         |                               |
| Austria               | .55                     | .43                           | .84                     | .69                           |
| Australia             | .95                     | .64                           |                         |                               |
| Canada                |                         |                               | .81                     | .63                           |
| Great Britain         | .84                     | .61                           | .99                     | .76                           |
| Italy                 | .74                     | .59                           | .91                     | .77                           |
| Netherlands           |                         |                               | .81                     | .61                           |
| New Zealand           | .85                     | .65                           | .59                     | .37                           |
| Norway                |                         |                               | .88                     | .68                           |
| Switzerland           | .84                     | .60                           | 1.19                    | .83                           |
| United States         | .84                     | .60                           | 1.05                    | .79                           |
| West Germany          | .84                     | .60                           |                         |                               |
| **Unweighted Mean**   | **.81**                 | **.60**                       | **1.90**                | **.68**                       |
| BIG G7 (US, WG, GB, It) | .85                    | .63                           | 1.04                    | .79                           |

Source: ISSP (1987, 1992). Calculated from data shown in Appendix Tables A1 and A2. For a list of the occupations covered, see text.

Perceived differentials increase between 1987 and 1992, and so too do “ought to” differentials. Countries with high perceived differentials have high “ought to” differentials.

To explore the relation between “ought to” and perceived occupational pay, we calculated individual-specific measures of perceived and “ought to” occupational earnings differentials. Specifically, for each person in the samples, we calculated the standard deviation of that respondent’s perception of actual earnings across occupations and the standard deviation of the earnings he or she believes “ought to” be paid across those occupations. These statistics differ from the statistics in Table 1. In Table 1 we calculated standard deviations for the natural log of mean earnings across occupations within a country (one value per country). Our new statistic is based on the standard deviation of the natural log of earnings across occupations reported by each individual (one value for each individual).

We then estimated ordinary least squares regressions of these individual measures of perceived and desired inequality on personal characteristics, such as schooling, age, gender, and country dummy variables, with the United States as the reference group. By including variables for individual characteristics, we perform *ceteris paribus* comparisons of people in the ex-communist countries and in the West and examine the relation between personal characteristics and perceived/"ought to" pay differentials. Since the group of countries covered in 1987 differs from that covered in 1992, we provide separate calculations for the two
years. To summarize the ex-communist countries/Western countries differentials of concern, we also estimated a restricted model in which we replaced the individual country dummy variable that takes the value 1 for an ex-communist country and 0 for the Western countries. Finally, to test inferences about changes over time, we further estimated our model on a pooled sample limited to countries that appear in both years.

Table 2 presents the basic cross-section regression results. There are three findings. First, consistent with Table 1, column 1 of Table 2 shows that, controlling for personal characteristics, in 1987 persons in ex-communist countries perceived lower occupational pay differentials than did respondents in the West, save for those in Australia. Column 2 shows that the Hungarians and Poles also believed that the differentials “ought to” be less than did respondents in the West, including the Australians. The coefficients on schooling further show that persons with more potential earnings power also perceived greater differentials and were more likely to believe that such differentials were fair.

Column 3 adds the standard deviation of the natural log of perceived differentials to the “ought to” equation. What does this tell us? It shows what people in different countries believe the spread in occupational pay should be, conditional on the spread they perceive. For instance, the coefficient on the dummy for Hungary of −0.0574 shows that, taking account of the lower perceived spread in earnings, Hungarians still desired a smaller spread than did persons in the excluded United States.

The coefficients on the standard deviation of perceived earnings are overwhelmingly statistically significant. This implies that people’s attitudes toward the spread of earnings that “ought to” be paid are formed not in a vacuum but in fact depend upon what they see in the marketplace. The 0.62 coefficient implies, roughly, that individuals believe the distribution of occupational differentials in their country “ought to” be scaled down from what they perceive it to be by a factor of 0.62.

For example, if the standard deviation of the natural log of perceived occupational earnings was 0.64 (the mean value in our sample in 1992), the estimated coefficient implies, ceteris paribus, individuals would like earnings differentials that would yield a standard deviation of 0.57 (the standard deviation of perceived pay of 0.92 multiplied by 0.62). Conditional on perceived differentials, persons in Hungary and Poland favored a modestly more egalitarian distribution than those in Western countries: the ex-communist dummy variable in the restricted regression is just −.02, though highly significant.

Columns 3–6 of the table confirm the more complicated pattern for 1992. Perceived differentials vary considerably, with dummies for the ex-communist countries showing that workers in most countries perceived less inequality than did persons in the United States (the big exception being Russia), but more than in Norway. The summary measure shows that the ex-communist dummy is a positive .10, due in large part to the low perceived inequality in Norway among the western countries. The regression for the spread in earnings that “ought to” be paid in column 5 shows a similar pattern, but in this case workers in all the ex-communist countries, including Russia, have lower “ought to” differentials than in the United States. The coefficient on the dummy variable for ex-communist countries in the summary regression is 0.0277.

This finding might lead some to believe that people in the ex-communist countries have less desire for narrowing pay differentials than do people in the West. But our hypothesis is that egalitarian desires are tempered by perceived market differentials. Thus the appropriate equation to assess the desires is equation (6). Here the dummy variables for the ex-communist countries are smaller than those in all Western countries except Norway, and the dummy variable in the summary regression is −0.029, with a t-statistic of 8.19. The similarity to the 1987 coefficient is impres-
Table 2. Perceptions of Actual Earnings and Fair Earnings: 
Regressions of Standard Deviations of Individuals’ Views, 1987 and 1992, in Fifteen Countries. 
(t-Statistics in Parentheses)

| Variable/Country          | Perceived Earnings 1987 | Earnings Viewed as Fair 1987 | Perceived Earnings 1992 | Earnings Viewed as Fair 1992 | Perceived Earnings 1992 | Earnings Viewed as Fair 1992 |
|---------------------------|-------------------------|-------------------------------|-------------------------|-------------------------------|-------------------------|-------------------------------|
|                           | (1)                     | (2)                           | (3)                     | (4)                           | (5)                     | (6)                           |
|                           |                         |                               |                         |                               |                         |                               |
| SD Perceived Earnings     | —                       | —                             | —                       | .6173 (74.49)                | —                       | —                             |
| Years of Schooling        | .0049 (6.69)            | .0070 (6.59)                  | .0038 (6.32)           | .0058 (8.02)                  | .0105 (15.25)          | .0071 (12.23)                 |
| Bulgarian                 | −.1967 (15.98)          | −.2176 (19.09)                | −.0741 (7.54)          | −.2220 (17.66)                | −.1905 (16.52)         | −.0271 (2.77)                 |
| Czechoslovakia            | −.1962 (15.64)          | −.1687 (14.86)                | −.0305 (3.03)          | −.2005 (15.97)                | −.1631 (14.48)         | −.0183 (1.87)                 |
| East Germany              | .0171 (1.41)            | −.0820 (7.24)                 | −.0630 (6.58)          | .0020 (0.17)                  | −.0612 (5.41)          | −.0314 (3.34)                 |
| Hungary                   | −.2591 (33.47)          | −.2209 (28.93)                | −.0574 (8.62)          | −.0466 (3.79)                 | −.1678 (15.18)         | −.1127 (11.53)                |
| Poland                    | −.2999 (35.22)          | −.2421 (28.99)                | −.0504 (6.84)          | −.1078 (9.09)                 | −.1366 (12.76)         | −.0503 (5.31)                 |
| Russia                    | .0695 (6.42)            | −.0338 (3.40)                 | −.0506 (5.86)          | .0576 (5.21)                  | −.183 (1.82)           | −.0183 (3.25)                 |
| Australia                 | −.3028 (35.79)          | −.2087 (24.14)                | −.0217 (2.95)          | −.2306 (22.28)                | −.1780 (18.55)         | −.0233 (2.76)                 |
| Austria                   | .0674 (6.95)            | −.0164 (1.71)                 | −.0495 (6.29)          | .0266 (2.23)                  | −.0392 (3.47)          | −.0022 (0.23)                 |
| Great Britain             | −.0850 (9.84)           | −.0956 (10.84)                | −.0439 (6.20)          | −.0266 (2.32)                 | −.0392 (3.47)          | −.0022 (0.23)                 |
| Italy                     | −.1026 (7.53)           | −.0200 (1.56)                 | .0575 (5.31)           | −.1154 (8.46)                 | −.1154 (1.32)          | −.0167 (7.06)                 |
| Netherlands               | −.1466 (17.84)          | −.1168 (13.93)                | −.0242 (3.53)          | −.2285 (20.07)                | −.1843 (17.13)         | −.0287 (3.09)                 |
| New Zealand               | —                       | —                             | —                       | −.2294 (20.10)                | −.1889 (17.70)         | −.0272 (3.01)                 |

Continued
Table 2. Continued

| Variable/Country | Perceived Earnings 1987 | Earnings Viewed as Fair 1992 | Perceived Earnings 1992 | Earnings Viewed as Fair 1992 | Perceived Earnings 1992 | Earnings Viewed as Fair 1992 |
|------------------|--------------------------|-------------------------------|--------------------------|-------------------------------|--------------------------|-------------------------------|
| Norway           | -0.4778                  | -0.3898                       | -0.1004                  | -0.4737                      | -0.3963                  | -0.1029                      |
|                  | (40.50)                  | (37.15)                       | (10.02)                  | (40.11)                      | (38.09)                  | (10.53)                      |
| Switzerland      | -0.1369                  | -0.1234                       | -0.0395                  | -0.1369                      | -0.1234                  | -0.0395                      |
|                  | (14.00)                  | (12.21)                       | (5.01)                   | (14.00)                      | (12.21)                  | (5.01)                       |
| West Germany     | -0.0803                  | -0.0841                       | -0.0461                  | -0.0266                      | -0.0249                  | -0.0136                      |
|                  | (9.34)                   | (9.73)                        | (6.56)                   | (2.23)                       | (2.53)                   | (1.61)                       |
| Income Differences Too High: |                      |                               |                          |                               |                          |                               |
| Strongly Disagree|                          |                               |                          |                               |                          |                               |
| Disagree         |                          |                               |                          |                               |                          |                               |
| Agree            |                          |                               |                          |                               |                          |                               |
| Strongly Agree   |                          |                               |                          |                               |                          |                               |
| R²               | 0.2923                   | 0.1933                        | 0.4930                   | 0.2706                       | 0.1945                   | 0.4788                       |
|                  | (6.49)                   | (5.32)                        | (19.63)                  | (14.76)                      | (8.59)                   | (1.56)                       |
| Root MSE         | 0.2004                   | 0.2012                        | 0.1594                   | 0.2406                       | 0.2387                   | 0.1855                       |
| (b) Summary Measure* |                |                               |                          |                               |                          |                               |
| Ex-Communist Dummy| -0.2089                  | -0.1578                       | -0.0223                  | 0.0707                       | -0.0289                  | 0.0764                       |
|                  | (42.71)                  | (20.23)                       | (5.32)                   | (19.63)                      | (6.19)                   | (8.19)                       |

*Here the coefficient and standard error are reported on a dummy in six different regressions in which ex-communist countries are set to one and all capitalist countries are set to zero.

Notes: Equations also include a constant, five age dummies, and a gender dummy. The excluded category is the United States. Because the schooling variable was missing in the case of Poland, we imputed the mean of the sample (11.15509) for all cases from that country.

Source: ISSP (1987, 1992).
sive, suggesting that the column (3) and (6) results represent a more stable pattern than the column (2) and (5) results.

How do the findings square with those reported earlier based on the broader question of the percentages who agree strongly that differences in income are too large? If the responses to the occupational earnings questions and the responses to the “Are differences in income too large?” question are capturing the same attitudes, we would expect that including the responses to the latter in the equations for the spread of “ought to” pay should greatly weaken the dummy variables for the ex-communist countries. Column 7 shows that this is indeed the case. The coefficients on virtually all of the country dummy variables decline, and the summary measure for ex-communist states drops to a bare -0.0055, which is not significantly different from zero. The underlying reason for the negative coefficient in column (6) is thus the broad desires for lower income differences.

Overall, the implication of Table 2 is that individuals in the ex-communist countries more strongly desired a narrowing of pay differentials than did those in the West, but that they substantially tempered that desire in the light of rising differentials during the transition to a market economy. The data on which the regressions reported in Table 2 are based, however, differ across the two years both in the mix of countries included and in the set of occupations for which earnings were reported. To make a firmer assessment of changes over time, and guarantee that any measured changes over time are not due to changes in the composition of our samples, we formed a pooled sample of observations for individuals in the ex-communist countries and for those in the Western countries that reported in both years. This sample, which contains just two ex-communist countries, Hungary and Poland, and just four Western countries, the United States, Great Britain, Australia, and West Germany, provides a relatively clean picture of changes over time. To capture the changes, we include dummy variables for interactions between country and year. The dummy variable for 1992 is included to capture actual changes across the years as well as differences in the samples of occupations between the two years.

Table 3 gives our pooled regression results. The dummy variables for Hungary and Poland in column 1 show that in 1987 perceived differentials were lower in those two countries than in the West. The 1992 country interaction terms show that the differentials rose by about 0.20 natural log points over the five-year period. The regression coefficients for the 1992 interactions also show rises in perceived differentials in Western countries, but by smaller magnitudes. To obtain the perceived differentials in 1992, we must sum the coefficients on country and country-year interactions. The results of that analysis show that perceived differentials in Hungary and Poland differed little from those in Great Britain and Germany, were smaller than in the United States, and were larger than in Australia.

The spread of “ought to” differentials in column 2 reveals a general upward trend in the differentials respondents believe “ought to” be paid: the 1992 interaction terms are positive. The upward trend is weaker in Hungary but not in Poland than in the Western countries: the coefficients on the country-year dummy variable of 0.0473 for Hungary suggest a smaller perceived increase than in the United States (the 1992 dummy of 0.0869 reflects the United States) or in the other Western countries, whereas for Poland, the 0.0997 coefficient is comparable to that for the United States and larger than those for other Western countries.

The bottom line calculations are in Table 3, where we estimate the coefficients on country dummy variables, conditional on the spread in perceived differentials. In 1987 people in both Poland and Hungary had a lower spread in “ought to” differentials than those in the United States or Australia. But their spread is similar to that among people in Great Britain and Germany. The sum of the coefficients on the country terms and the country-year interaction terms for 1992 shows that the spread in pay across occupations that was believed to
Table 3. Perceptions of Actual Earnings and Fair Earnings in Five Countries: Pooled Regressions, 1987 and 1992. (t-Statistics in Parentheses)

| Variable/Country          | Perceived Earnings | Earnings Viewed as Fair | Earnings Viewed as Fair |
|---------------------------|--------------------|-------------------------|-------------------------|
|                           | (1)                | (2)                     | (3)                     |
| SD Perceived Earnings     | —                  | —                       | .5992                   |
|                           | (86.28)            | (9.92)                  |                         |
| Years of Schooling        | .0068              | .0101                   | .0057                   |
|                           | (9.02)             | (12.49)                 | (9.43)                  |
| Hungary                   | −.2556             | −.2156                  | −.0592                  |
|                           | (30.89)            | (26.31)                 | (8.43)                  |
| Poland                    | −.2973             | −.2380                  | −.0530                  |
|                           | (32.73)            | (26.60)                 | (6.85)                  |
| Australia                 | −.2992             | −.2019                  | −.0225                  |
|                           | (32.25)            | (21.86)                 | (2.93)                  |
| Great Britain             | −.0835             | −.0924                  | −.0431                  |
|                           | (9.03)             | (9.76)                  | (5.65)                  |
| West Germany              | −.0765             | −.0778                  | −.0432                  |
|                           | (8.36)             | (8.42)                  | (5.73)                  |
| 1992 Dummy                | .1466              | .0869                   | −.0248                  |
|                           | (14.99)            | (9.13)                  | (3.02)                  |
| Hungary*1992              | .2115              | .0473                   | −.0547                  |
|                           | (15.51)            | (3.65)                  | (4.83)                  |
| Poland*1992               | .1922              | .0997                   | .0011                   |
|                           | (13.97)            | (7.57)                  | (0.09)                  |
| Australia*1992            | .0688              | .0210                   | .0004                   |
|                           | (5.42)             | (1.66)                  | (0.04)                  |
| Great Britain*1992        | .0571              | .0506                   | .0379                   |
|                           | (4.07)             | (3.60)                  | (3.29)                  |
| West Germany*1992         | .0579              | .0517                   | .0533                   |
|                           | (4.50)             | (4.07)                  | (5.01)                  |

Notes: Equations also include five age dummies and a gender dummy. The excluded category is the United States. Because the schooling variable was missing for Poland, in all cases for that country we imputed the mean for the other countries.

Source: ISSP (1987, 1992).

be "fair" was smaller for Hungarians and Poles than for Americans or other Westerners, conditional on the perceived spread.

We conclude that respondents in the ex-communist countries had a greater preference for narrowing occupational earnings, conditional on perceived earnings, than did Westerners, but that the difference in preferences diminished greatly as perceived differentials widened in those countries more rapidly than they did in most Western countries.

Job Satisfaction, Happiness and the Role of Trade Unions

Because Hungary has been in the ISSP since 1986, we have available more information on outcomes and attitudes for that country than for the other ex-communist
countries. The 1989 ISSP asked questions about job satisfaction and working conditions. The 1990 ISSP asked questions about unionism. Although we are leery about generalizing from a single country to all ex-communist countries, the differences in attitudes between Hungarian employees and those in Western countries do provide insight into the potential legacy of communism. They allow us to examine the widely held view that communist labor relations practices produced less desirable workplace conditions and less satisfied workers than free market practices. (See Haraszti 1978; for a contrasting view, however, see Burawoy and Lukacs 1992.)

Respondents to the 1989 ISSP were asked the generic (and widely used) job satisfaction question, “How satisfied are you in your job?” Questions about job satisfaction are difficult to interpret due to the subjective nature of the variable and problems in making interpersonal comparisons (Freeman 1978). Still, econometric analyses based on satisfaction data have yielded interesting and consistent results across data sets that show links between satisfaction and economic and demographic variables, and also show that job satisfaction is a good predictor of future quit behavior—indeed, in many longitudinal surveys, the best such predictor (see, for example, Hamermesh 1977; Borjas 1979; Freeman 1978; Blanchflower and Oswald 1992; and Clark and Oswald 1997).

What is more troubling for our analysis is that people in one country may “scale” responses differently from those in another (see Hofstede 1985 for studies of different country responses to satisfaction-type questions in a single multinational). For instance, Americans may be relatively optimistic, with an “everything will work out” mentality that leads them to respond more positively than a comparable British group to the question “Are you satisfied with your job?” even though their true satisfaction, on some objective scale, is the same as that of the (possibly more reserved) British. We deal with this problem by examining answers to more objective questions regarding workplace conditions—such as health and safety conditions—and estimating an ordered probit model in which we include as controls responses to other unrelated questions that might index national patterns of response.

In any case, relatively few Hungarians reported being very satisfied or completely satisfied (13%) with their jobs compared to Westerners, among whom the proportion ranged from 33% in Eire to 50% in the United States. At face value, this confirms the notion that workers under communism are less satisfied than workers under capitalism.

Table 4 presents evidence on worker perceptions of objective workplace conditions that might explain the low job satisfaction reported by Hungarian workers. It shows that Hungarian workers are far less likely to regard their jobs as interesting than are Western workers (question 1) and far more likely to see their workplace as involving dangerous conditions (question 2) and unhealthy conditions (question 3). All of these perceptions are likely to feed into low job satisfaction. To the extent that they reflect objective conditions, they suggest that the reported lower job satisfaction has some grounding in reality.

As our final check on the effect of communist labor relations on job satisfaction, we use the entire distribution of responses on the 1989 job satisfaction question to estimate the difference in job satisfaction between Hungarian workers and Western

Satisfaction correlates well with quit behavior, so if the same response led to similar quitting across countries, we could view responses as valid indicators of one form of objective behavior. However, we do not have data contrasting satisfaction and quits across countries.

The proportions of workers who reported being "very" or "completely" satisfied with their jobs were as follows: Hungary, 13%; Ireland, 33%; Italy, 34%; Israel, 37%; Great Britain, 39%; Netherlands, 39%; Norway, 43%; West Germany, 44%; Austria, 48%; and the United States, 50%.

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11One objective scaling would be to determine how satisfaction responses map into labor turnover.
is negatively correlated with satisfaction (Freeman 1978; Borjas 1979), and self-employment, which has been found to be positively related to satisfaction (Blanchflower and Oswald 1993). In this analysis, the coefficient on this dummy variable tells us how the entire distribution of job satisfaction differs between Hungarians and Westerners, conditional on the diverse control variables.

Table 5 presents our results. Column 1 gives a basic job satisfaction equation. It shows that the “control variables” have their expected effects on job satisfaction. Thus, the ISSP question on job satisfaction gives results comparable to those in other surveys: unionism, in particular, is negatively related to job satisfaction, while self-employment is positively related. The coefficient on the Hungary dummy is large, negative, and statistically significant (–.53).

Column 2 controls, as best we can, for income, on the hypothesis that Hungarians may be less satisfied with their job not because of any legacy of communist labor relations but simply because they are lower paid than Western workers. For this purpose, we use an ISSP question that asked people whether they viewed their income as high: “For each of the following statements about your job, please tick one box to show how much you agree or disagree that it applies to your job: My income is high.”

Consistent with lower income in former communist countries, Hungarians were far more likely than Western workers to disagree with this statement. Some 70% of Hungarians disagreed or disagreed strongly that their income was high, compared to 48% of the British, 27% of Germans, 33% of Italians, 43% of Americans, and comparable fractions of other Westerners. Workers who report that their income is low are less likely to be satisfied (that is, the coefficients on the various categorical responses are increasingly negative compared to the omitted group of persons who strongly agree that their income is high). The coefficient on the Hungary dummy falls to –.41 upon addition of the “income is high” responses,

workers in an ordered probit model. Ordered probits are the appropriate statistical procedure where, as in this case, respondents express their preferences in the form of an ordinal ranking.

Our model includes various demographic variables that are known to influence job satisfaction, including gender, age, and marital status. We also include a unionism variable, which previous studies have found

Table 4. Employees’ Perceptions of
Their Work and Working Conditions, 1989:
Respondents in Hungary Compared
to Those in Eleven Capitalist Countries.a

| Response                  | Hungary |
|---------------------------|---------|
| Strongly Agree            | 6.6     |
| Agree                     | 60.4    |
| Neither                   | 20.4    |
| Disagree                  | 10.8    |
| Strongly Disagree         | 1.5     |
| N                         | 603     |
| (Q1) “My job is interesting” - Employed Respondents Only (1989: V63) |
| Always                    | 15.8    |
| Often                     | 13.0    |
| Sometimes                 | 16.3    |
| Hardly Ever               | 10.8    |
| Never                     | 44.2    |
| N                         | 602     |

*For a list of countries, see Table 1 or text. Source: ISSP (1989).*

| Always                    | 10.3    |
| Often                     | 9.8     |
| Sometimes                 | 17.6    |
| Hardly Ever               | 12.4    |
| Never                     | 49.9    |
| N                         | 602     |

(Q2) “How often do you work in dangerous conditions?” - Employed Respondents Only (1989: V72)

(Q3) “How often do you work in unhealthy conditions?” - Employed Respondents Only (1989: V73)
Table 5. Correlates of Job Satisfaction, 1989: Ordered Probits.

| Variable                  | (1)       |                     | (2)       |                     | (3)       |                     |
|---------------------------|-----------|---------------------|-----------|---------------------|-----------|---------------------|
|                           | Coefficient | Standard Error | Coefficient | Standard Error | Coefficient | Standard Error |
| Hungary Dummy             | -0.5315   | 0.0471              | -0.4149   | 0.0479              | -0.3211   | 0.0500              |
| Male                      | -0.1207   | 0.0271              | -0.2015   | 0.0277              | -0.1271   | 0.0296              |
| Self-Employed             | 0.3242    | 0.0468              | 0.3249    | 0.0476              | 0.2334    | 0.0487              |
| Married                   | 0.0207    | 0.0298              | 0.0011    | 0.0302              | -0.0190   | 0.0308              |
| Age                       | 0.0100    | 0.011               | 0.0091    | 0.0012              | 0.0066    | 0.0012              |
| Years of Schooling        | 0.0002    | 0.0044              | -0.0089   | 0.0045              | -0.0356   | 0.0047              |
| Union Status              | -0.0958   | 0.0279              | -0.1030   | 0.0281              | -0.0708   | 0.0288              |
| "My income is high"a      |           |                     |           |                     |           |                     |
| Agree                     | -0.2247   | 0.0813              | -1.495    | 0.0832              |           |                     |
| Neither Agree Nor Disagree| -0.5683   | 0.0791              | -3.626    | 0.0814              |           |                     |
| Disagree                  | -0.8650   | 0.0792              | -5.869    | 0.0917              |           |                     |
| Strongly Disagree         | -1.2801   | 0.0872              | -9.870    | 0.0895              |           |                     |
| Can’t Choose              | -0.4569   | 0.1641              | -1.730    | 0.1730              |           |                     |
| "My job is interesting"a  |           |                     |           |                     |           |                     |
| Agree                     | -0.5806   | 0.0535              |           |                     |           |                     |
| Neither Agree Nor Disagree| -1.2294   | 0.0482              |           |                     |           |                     |
| Disagree                  | -1.5457   | 0.0621              |           |                     |           |                     |
| Strongly Disagree         | -2.1059   | 0.0978              |           |                     |           |                     |
| Can’t Choose              | -1.0886   | 0.2808              |           |                     |           |                     |
| "Dangerous conditions"    |           |                     |           |                     |           |                     |
| Often                     | -0.1899   | 0.0848              |           |                     |           |                     |
| Sometimes                 | -0.0442   | 0.0764              |           |                     |           |                     |
| Hardly Ever               | -0.1519   | 0.0777              |           |                     |           |                     |
| Never                     | -1.096    | 0.0752              |           |                     |           |                     |
| "Unhealthy conditions"b   |           |                     |           |                     |           |                     |
| Often                     | 0.0647    | 0.0868              |           |                     |           |                     |
| Sometimes                 | 0.1185    | 0.0775              |           |                     |           |                     |
| Hardly Ever               | 0.2159    | 0.0788              |           |                     |           |                     |
| Never                     | 0.4109    | 0.0756              |           |                     |           |                     |
| Threshold 1               | -2.2709   | 0.0911              | -3.2536   | 0.1238              | -4.2908   | 0.1553              |
| Threshold 2               | -1.9239   | 0.0818              | -2.8934   | 0.1165              | -3.8575   | 0.1469              |
| Threshold 3               | -1.3500   | 0.0759              | -2.2885   | 0.1117              | -3.1390   | 0.1411              |
| Threshold 4               | -0.7704   | 0.0741              | -1.6779   | 0.1100              | -2.4310   | 0.1391              |
| Threshold 5               | 0.5209    | 0.0739              | -0.3192   | 0.1089              | -0.8877   | 0.1375              |
| Threshold 6               | 1.4285    | 0.0751              | 0.6355    | 0.1091              | 0.1621    | 0.1370              |
| Number of Observations    | 6399      |                     | 6346      |                     | 6312      |                     |
| Chi²(7)                   | 307.40    |                     | 844.94    |                     | 2133.24   |                     |
| Pseudo R²                 | 0.0170    |                     | 0.0471    |                     | 0.1194    |                     |
| Log Likelihood            | -8907.452 |                     | -8555.3416|                     | -7866.71  |                     |

Note: The sample consists of employed respondents only.
Source: ISSP (1989).
"The excluded category is “strongly agree.”
"The excluded category is “always.”
but it remains substantial and significant. At similar perceptions of income (if not similar incomes), Hungarians are less satisfied with their jobs than Westerners. Finally, column 3 includes a series of dummy variables to distinguish whether respondents considered their jobs to be “interesting” or their working conditions to be “unhealthy” or “dangerous” (as in Table 4). The inclusion of these variables reduces the size of the coefficient on the Hungary dummy by approximately a quarter, although the coefficient remains statistically significant. These factors contribute to the low levels of satisfaction in Hungary, but they are not the whole story. All told, we interpret the ordered probits as confirming the reported lower job satisfaction of Hungarians with evidence about objective features of workplaces.

In addition to evidence on job satisfaction, we have reports of overall happiness by respondents, whether employed, unemployed, or out of the labor force. Since work is such an important part of life, we expect that this measure of attitudes ought also to be related to labor market experiences. Our “happiness” measure also has one important advantage over job satisfaction: it covers all citizens, including the jobless. The 1991 ISSP provides data on perceived “happiness” for Slovenia, Poland, and East Germany, as well as for Hungary and Western countries. Respondents were asked: “If you were to consider your life in general these days, how happy or unhappy would you say you are, on the whole?” Four options were given for their replies—“not at all,” “not very,” “fairly,” and “very.” Thirty-eight percent of Hungarians, 23% of East Germans, and 45% of Slovenes said they were not at all happy or not very happy, compared to 11% of West Germans, 8% of the British, 8% of Americans, 7% of the Irish, 10% of New Zealanders, and 11% of Norwegians. Only the Italians and Israelis had “not at all” or “not very” happy percentages close to those of the eastern Europeans—22% for Italians and Israelis.

Table 6 reports a series of ordered probit equations to estimate country effects on happiness, controlling for age, gender, marital status, and years of schooling. Once again we report results including a full set of country dummies (the excluded country is the United States) as well as the results from an equation that replaces the country dummies with a single ex-communist dummy. Individuals in the ex-communist countries reported being less happy than those in the Western countries in our sample. As can be seen from column 1, the lowest levels of happiness among the Western countries were to be found in Israel and Italy. The young, married individuals, women, and those with higher levels of schooling reported the highest levels of happiness.

In order to obtain results comparable to the job satisfaction evidence presented above, which was for workers only, in columns 2 and 3 we restrict the samples to the employed and those unemployed or out of the labor force, respectively. The pattern of results, including the country dummy structure, is remarkably similar to the job satisfaction pattern. In column 4, where two dummies are added to control for labor market status, we confirm the earlier finding of Oswald (1994) that the unemployed are less happy than the employed (the excluded category). This negative unemployment result is further confirmed in columns 5 and 6 of Table 6, where we report separate estimates for the 11 capitalist countries and the 4 ex-communist countries. The main differences between these two equations are the lack of significance of the

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13 Overall, the responses were as follows (%):

|          | Not at All | Not Very | Fairly | Very | No. Obs. |
|----------|------------|----------|--------|------|----------|
| All      | 3          | 15       | 60     | 22   | 17967    |
| Capitalist | 2          | 9        | 62     | 27   | 12832    |
| Ex-Communist | 5          | 30       | 56     | 9    | 5135     |

14 The excluded countries are the United States and Hungary, respectively.
Table 6. Reported Happiness in Capitalist and Ex-Communist Countries, 1991: Ordered Probits.

| Variable/Country         | All (1) | Employed (2) | Not Employed (3) | All (4) | Capitalist (5) | Ex-Communist (6) |
|--------------------------|---------|--------------|------------------|---------|----------------|------------------|
| Male                     | -0.0933 | -0.0980      | -0.1418          | -0.0960 | -0.1232        | -0.0403          |
| Age 25–34                | -0.1901 | -0.1748      | -0.2785          | -0.2068 | -0.2025        | -0.2181          |
| Age 35–44                | -0.2742 | -0.2888      | -0.3113          | -0.3003 | -0.2844        | -0.3406          |
| Age 45–54                | -0.2727 | -0.2930      | -0.2774          | -0.3037 | -0.2898        | -0.3376          |
| Age 55–64                | -0.2738 | -0.2823      | -0.2277          | -0.2994 | -0.2921        | -0.3087          |
| Age 65–74                | -0.1158 | 0.1367       | -0.0869          | -0.1373 | -0.0839        | -0.2994          |
| Married                  | 0.3659  | 0.3532       | 0.3816           | 0.3547  | 0.3670         | 0.3004           |
| West Germany             | -0.3661 | -0.3574      | -0.3819          | -0.3748 | -0.3904        | n/a              |
| Great Britain            | -0.1231 | -0.1412      | -0.1031          | -0.0980 | -0.1065        | n/a              |
| Hungary                  | -1.0946 | -1.0500      | -1.2126          | -1.1122 | n/a            | n/a              |
| Netherlands              | -0.2018 | -0.2147      | -0.1755          | -0.1961 | -0.2006        | n/a              |
| Italy                    | -0.8122 | -0.8222      | -0.8047          | -0.8143 | -0.8234        | n/a              |
| Eire                     | 0.0863  | 0.1638       | 0.1049           | 0.1165  | 0.1127         | n/a              |
| Northern Ireland         | -0.0090 | 0.1081       | -0.1142          | 0.0111  | 0.0025         | n/a              |
| Poland                   | -0.9233 | -0.8681      | -0.9875          | -0.9003 | n/a            | 0.1752           |
| Slovenia                 | -1.2245 | -1.2372      | -1.2220          | -1.2226 | n/a            | -1.301           |
| Philippines              | -0.4325 | -0.4144      | -0.4701          | -0.4343 | -0.4492        | n/a              |
| New Zealand              | -0.2824 | -0.2678      | -0.2986          | -0.2836 | n/a            | -0.2884          |
| Norway                   | -0.4417 | -0.4662      | -0.3958          | -0.4459 | n/a            | -0.4537          |
| Israel                   | -0.8190 | -0.6870      | -0.9931          | -0.7888 | n/a            | -0.7774          |
| East Germany             | -0.8285 | -0.8425      | -0.8258          | -0.8114 | n/a            | 0.2898           |

Continued
gender variable in equation 6 and the higher levels of happiness of those age 65–74 in equation 5 compared with the middle-aged. Individuals living in ex-communist countries report lower levels of happiness and, if employed, lower levels of job satisfaction than are found in our control group of Western countries. This is further confirmation of the legacy hypothesis.

**Attitudes Toward Unions**

One interpretation of the results on job satisfaction is that it is a legacy of transmission belt unions and the absence of a free labor market. If the lack of independent voice is, in fact, an important factor, we would expect to see Hungarians and persons in other ex-communist countries differ from people in Western countries in their attitudes toward trade unions.

The 1990 and 1993 ISSP contains several questions about attitudes toward unionism that enable us to test that interpretation. Table 7 records the responses of Hungarians, Bulgarians, and Czechs as well as ISSP respondents from Western countries to questions regarding unionism in those surveys. The responses to the question “How good are trade/labor unions for the country as a whole?” show that persons in the ex-communist countries had less favorable views of unions in their country than did Westerners. Forty-two percent of Hungarians, 45% of Bulgarians, and 37% of Czechs reported that unions were “not very good” or “not good at all,” compared to 16% of West Germans, 27% of the British respondents, 25% of American respondents, and so on.

A large number of Italians (44%) also reported that unions were “not very good” or “not good at all.” But the reason for the dissatisfaction with unions in ex-communist countries differs markedly from the reason for Italian dissatisfaction. Asked “Do you think that trade unions have too much power or too little power?”, the eastern Europeans overwhelmingly thought that unions had too little power, whereas Italians disapproved of unions as having too much power (responses to question 2). Moreover, the proportion of Hungarians who expressed the belief that strong trade unions are needed to protect workers (question 3) exceeded
Table 7. Perceptions of the Role of Trade Unions in Nine Countries, 1990 and 1993.

(Q1) “How good are trade/labor unions for the country as a whole?”  
(1990: V45 and 1993 preliminary for Bulgaria & Czech Republic)  

| Response          | Hungary 1990 | Bulgaria 1993 | Czech Republic 1993 | Ex-Communist 1990-1993 | Capitalist 1990 | West Germany 1990 | GB 1990 | USA 1990 | Norway 1990 | Israel 1990 | Italy 1990 |
|-------------------|--------------|---------------|---------------------|------------------------|----------------|-------------------|--------|---------|-------------|-------------|-----------|
| Excellent         | 2.7          | 10.1          | 4.0                 | 5.6                    | 3.3            | 5.7               | 2.0    | 2.9     | 3.7          | 3.3          | 1.9       |
| Very Good         | 30.4         | 21.4          | 15.6                | 22.5                   | 17.3           | 26.0              | 15.7   | 17.2    | 21.9         | 15.2         | 7.9       |
| Fairly Good       | 25.1         | 23.1          | 57.6                | 35.3                   | 53.5           | 53.3              | 55.3   | 54.3    | 58.6         | 53.1         | 46.5      |
| Not Very Good     | 26.6         | 21.2          | 18.8                | 22.2                   | 19.9           | 13.2              | 20.7   | 19.7    | 12.2         | 21.3         | 32.4      |
| Not Good at All   | 15.2         | 24.2          | 3.9                 | 14.4                   | 6.0            | 1.8               | 6.2    | 5.7     | 3.6          | 7.1          | 11.3      |
| N                 | 782          | 822           | 865                 | 2469                   | 7189           | 2303              | 971    | 943     | 1271         | 750          | 951       |

(Q2) “Do you think that trade/labor unions in this country have too much power?”  
(1990: V42 and 1993 preliminary for Bulgaria & Czech Republic)  

| Response          | Hungary 1990 | Bulgaria 1993 | Czech Republic 1993 | Ex-Communist 1990-1993 | Capitalist 1990 | West Germany 1990 | GB 1990 | USA 1990 | Norway 1990 | Israel 1990 | Italy 1990 |
|-------------------|--------------|---------------|---------------------|------------------------|----------------|-------------------|--------|---------|-------------|-------------|-----------|
| Far Too Much      | 5.8          | 5.1           | 1.4                 | 4.1                    | 13.4           | 8.4               | 9.1    | 14.7    | 12.3         | 7.8          | 28.2      |
| Too Much Power    | 12.4         | 19.9          | 6.8                 | 13.0                   | 25.4           | 19.5              | 26.8   | 32.0    | 24.9         | 30.4         | 18.9      |
| About Right       | 18.3         | 20.7          | 26.7                | 21.9                   | 44.0           | 57.9              | 49.3   | 39.3    | 52.2         | 33.2         | 32.2      |
| Too Little Power  | 33.7         | 32.6          | 50.2                | 38.8                   | 14.3           | 12.3              | 13.0   | 11.5    | 9.5          | 25.3         | 14.3      |
| Far Too Little    | 29.8         | 21.7          | 14.8                | 22.1                   | 2.8            | 1.8               | 1.7    | 2.4     | 1.1          | 3.2          | 6.4       |
| N                 | 782          | 792           | 849                 | 2423                   | 7195           | 2303              | 971    | 943     | 1269         | 779          | 930       |

(Q3) “Workers need strong trade unions to protect their interests” (1989: V23)  

| Response           | Hungary Unweighted Mean | West Germany | GB | USA | Austria | Norway | Netherlands | Israel | Italy |
|--------------------|-------------------------|--------------|----|-----|---------|--------|-------------|--------|-------|
| Strongly Agree     | 28.8                    | 21.3         | 31.5 | 12.1 | 10.6    | 24.4   | 23.7        | 17.7   | 24.2  |
| Agree              | 47.4                    | 39.8         | 41.7 | 29.6 | 25.5    | 49.0   | 41.2        | 48.2   | 42.8  |
| Neither            | 12.0                    | 18.8         | 15.4 | 23.2 | 26.2    | 11.0   | 19.2        | 22.5   | 16.2  |
| Disagree           | 10.0                    | 15.0         | 9.1  | 28.8 | 25.4    | 12.2   | 11.1        | 9.1    | 12.2  |
| Strongly Disagree  | 1.8                     | 5.1          | 2.3  | 6.3  | 12.3    | 3.5    | 4.8         | 2.6    | 4.3   |
| N                  | 880                     | 18273        | 1288 | 1085 | 1231    | 1631   | 1652        | 1487   | 1003 |

Source: ISSP (1990, 1993).

that in all Western countries except Germany.\(^{15}\)

\(^{15}\)We initially expected one additional ISSP question—"In general, how would you describe relations at your workplace between management and employees?"—to be informative on possible attitudes toward the need for unions. But the response of Hungarians relative to Westerners here was ambiguous. A much smaller proportion (11.5%) stated that labor-management relations were very good than in any Western country (the Netherlands was the next lowest at 18.5%), but at the same time, the proportion of Hungarians describing labor relations as quite or very bad (31.4%) was much smaller than the proportion of workers in Western countries (the next lowest was Israel, with 4%).
tries that unions are not good for the country, are too weak, and are needed to protect workers? Our explanation is that these responses reflect two aspects of the experience with unions under communism: the past role of unions as transmission belts of the state, and the weakness of newly emerging or changing traditional unions—one of several legacies of the country's communist labor system.

**Attitudes Toward the State in Economic Life**

Under communism, the state dominated economic life, with adverse consequences for economic progress and for worker and citizen satisfaction. By overturning communism, these countries have committed themselves to developing a market economy with a much smaller governmental role in economic affairs than in their past. Still, the decades of government-dominated economic life and communist egalitarian ideology may have left a legacy of "statism" in these countries that would produce attitudes toward state interventions in wage and employment different from those found among people brought up in market economies.

Table 8 presents responses to ISSP questions from the 1987, 1990, 1992, and 1993 modules that cast light on this issue. It shows a wide difference in attitudes toward state interventions between eastern European and Western countries, and some differences among the ex-communist countries as well. The figures under the heading

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*Source: ISSP (1987, 1990, 1992, 1993).*

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**Table 8. Percentages Who Strongly Favor Government Intervention in the Labor Market: Seven Ex-Communist Countries and Fifteen Capitalist Countries, Various Years.**

| Country          | Reduce Income Differences | Control Wages by Law | Provide Jobs for All | Provide Basic Income for All |
|------------------|---------------------------|-----------------------|----------------------|-----------------------------|
|                  | 1987 1992 1993            | 1990                  | 1987 1992            | 1987 1992                   |
| Bulgaria         | 61 63 70                  | 36                    | 30                   | 40 51                       |
| Czechoslovakia   | 23 34                     | 50                    | 55                   | 50                          |
| East Germany     | 42 39 50                  | 45                    | 52                   | 40                          |
| Hungary          | 32 33 38                  | 16                    | 72 50                | 47                          |
| Poland           | 36 30 38                  | 72                    | 50                   | 47                          |
| Russia           | 39 32                     | 75                    | 71                   |                             |
| Slovenia         | 31 32                     | 36                    | 27                   |                             |

*Unweighted Mean*: 34 33 39 33 62 51 40 50

| Capitalist       | Reduce Income Differences | Control Wages by Law | Provide Jobs for All | Provide Basic Income for All |
|------------------|---------------------------|-----------------------|----------------------|-----------------------------|
| Australia        | 10                        | 10                    | 10                   | 10                          |
| Austria          | 35 27                     | 36 31                 | 20 13                |                             |
| Canada           | 17 12                     | 16                    | 19                   |                             |
| Great Britain    | 21 27 22                  | 6                     | 24 21                | 20 28                       |
| Israel           | 32 24                     | 21                    | 37                   |                             |
| Italy            | 40 29                     | 21                    | 57                   |                             |
| Japan            | 30                        |                       |                      |                             |
| Netherlands      | 17 13                     | 23                    | 10                   |                             |
| New Zealand      | 16 15                     |                       | 19                   | 21                          |
| Norway           | 17 15 7                   |                       | 37                   | 35                          |
| Spain            | 19                        |                       |                      |                             |
| Sweden           | 16                        |                       | 26                   | 9                           |
| Switzerland      | 11                        | 17                    | 12                   |                             |
| USA              | 8 10 7                    | 7                     | 17 17                | 7 10                        |
| West Germany     | 17 20 23                  | 9                     | 35 22                | 17 19                       |

*Unweighted Mean*: 18 20 20 12 25 26 14 20
“Reduce Income Differences” show that a larger proportion of workers in the former communist countries than in the capitalist countries believe that the government is responsible for reducing differences in income. In 1987, 34% of persons in the two ex-communist countries for which we have data, Hungary and Poland, strongly agreed with this statement, proportions that far exceed those in Western countries save for Austria. In 1992, for which we have information for seven ex-communist countries, the results are similar. On average, 33% of people in the ex-communist countries agreed strongly that governments should reduce income differences, compared to just 20% in the capitalist countries (though many in the capitalist countries agreed somewhat). The sole countries whose responses overlap those in the other set of countries in this contrast are Italy and Austria among the Western countries and Czechoslovakia among the ex-communist countries. The figures for 1993 show an even greater divergence between the ex-communist and the capitalist countries, with no overlap at all!

There is less consistency among the former communist countries in whether respondents support controlling wages by law. The East Germans strongly favor such controls, but the Hungarians do not. The contrast between the East and West Germans on this issue is striking.

On the employment side, the ISSP asked in 1987 and in 1992 whether the government should provide jobs for all. The respondents in the two eastern European countries covered in 1987, Hungary and Poland, were far more likely to agree with this proposition than were respondents in the Western countries. In 1992, when the ISSP asked this same question in six ex-communist countries, the unweighted average proportion of persons in those countries who agreed strongly was 51%, compared to 26% in the capitalist countries.

The greater support for government playing a role in income and employment determination in former communist countries shows up, finally, in responses to the following question: “Should government provide basic income for all?” Only Hungary provided data on this question in 1987. There, the proportion strongly agreeing was 40%, far above the corresponding proportion in capitalist countries. The difference in 1992, for which we have data for seven ex-communist countries, is even greater.

In sum, while there are some country differences, the overwhelming pattern in these data is a greater proclivity for relying on the state among persons from ex-communist countries than among Westerners—consistent with the legacy hypothesis.

**Conclusions**

We have uncovered substantial differences between survey respondents in former communist countries and those in Western countries in responses to diverse questions about attitudes toward earnings inequality, job satisfaction, working conditions, unions, and the role of the state in regulating labor market outcomes. In ISSP surveys conducted over the years 1987–93, the citizens of former communist countries evinced a greater desire for egalitarianism, less satisfaction with their jobs, more support for strong unions, and more support for state intervention than did most Westerners. At the same time, as perceived occupational earnings differentials widened in these countries over the course of the breakdown of communism and the transition to more capitalistic systems between 1987 and 1992, tolerance for those differentials increased substantially. Without gainsaying the important differences in outcomes and attitudes among ex-communist countries and among Western countries, we believe that it is reasonable to interpret the broad pattern of differences between the two groups as reflecting a legacy of communist economics.

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16Frentzel-Zagorska and Zagorski’s (1993) analysis of Polish public opinion polls gave results that are consistent with ours. Table 7 in Rose and Haerpfer (1994) also shows substantial “collectivist values” in eastern European countries, though with variation among countries and questions.
Appendix Table A1
Responses to the ISSP, 1987–1992

| Country       | 1987 | 1989 | 1990 | 1991 | 1992 | 1993 | All  |
|---------------|------|------|------|------|------|------|------|
| Ex-Communist: |      |      |      |      |      |      |      |
| Bulgaria      | 1198 | 1183 | 2381 |      |      |      |      |
| Czechoslovakia| 1101 | 1005 | 2106 |      |      |      |      |
| East Germany  |      |      |      | 1028 | 1486 | 1094 | 1092 |
| Hungary       | 2606 | 1000 | 977  | 1000 | 1250 | 1167 | 8000 |
| Poland        | 3943 |      |      | 1063 | 1636 | 1641 | 8283 |
| Russia        |      |      |      |      |      | 1983 | 1931 |
| Slovenia      | 2080 | 1049 | 1032 |      |      |      | 4161 |
| Capitalist:   |      |      |      |      |      |      |      |
| Australia     | 1574 | 2398 |      | 2203 |      |      | 6175 |
| Austria       | 972  | 1997 |      | 1027 |      |      | 3996 |
| Canada        |      |      |      |      | 1004 | 1467 | 2471 |
| Eire          |      |      | 972  | 1005 | 1005 | 957  | 3939 |
| Great Britain |      |      |      |      |      |      |      |
| Israel        | 1212 | 1297 | 1197 | 1257 | 1066 | 1261 | 7290 |
| Italy         |      |      |      |      |      |      |      |
| Netherlands   |      |      |      |      |      |      |      |
| New Zealand   |      |      |      |      |      |      |      |
| Norway        |      |      |      |      |      |      |      |
| Sweden        | 1848 | 1517 | 1506 | 1538 | 1414 |      | 7823 |
| Switzerland   |      |      |      |      |      |      | 987  |
| USA           | 1564 | 1453 | 1217 | 1359 | 1273 | 1557 | 8423 |
| West Germany  | 1397 | 1575 | 2812 | 1346 | 2297 | 1014 | 10441|
| Total         | 16920| 16391| 11727| 16781| 22703| 25969| 107978|

**Notes:** Data for 1993 were obtained from the ISSP 1993 Source book (ZA Study 2450). Data are also available in the ISSP (but not used in this paper) for Northern Ireland and the Philippines for a number of years and Spain and Japan for 1993 only.

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