Diagnostic value of soluble mesothelin-related peptides in pleural effusion for malignant pleural mesothelioma

An updated meta-analysis

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Abstract

Background: Soluble mesothelin-related peptide (SMRP) is a widely studied tumor marker for diagnosing malignant pleural mesothelioma (MPM). This study discussed the diagnostic value of SMRPs in pleural effusion (PE) for MPM.

Methods: Medline, Embase, Web of Science, and Cochrane library system were systematically searched on the data of SMRPs in PE for MPM diagnosis. Pooled diagnostic sensitivity, specificity, and symmetric receiver operating characteristic curve were calculated.

Results: Thirteen studies fulfilled the inclusion criteria and a total of 3359 cases including 759 MPM cases, 1061 non-MM (malignant mesothelioma) malignant PE, and 1539 benign PE were brought into this meta-analysis. The pooled results of SMRPs in PE for diagnosing MPM were as follows: sensitivity, specificity, positive likelihood ratio, negative likelihood ratio, and diagnostic odds ratio were 0.68 (95% confidence interval [CI]: 0.64–0.72), 0.91 (95% CI: 0.86–0.94), 7.8 (95% CI: 5.0–12.0), 0.35 (95% CI: 0.31–0.40), and 22 (95% CI: 14–35), respectively. The area under the summary receiver operating characteristic curves (AUC) was 0.75 (95% CI: 0.72–0.80). Subgroup analyzes revealed that the AUC of cohort group using histological diagnosis could be improved to 0.86 (95% CI: 0.83, 0.89). The Deek’s funnel plot asymmetry test showed no publication bias.

Conclusion: Although the sensitivity of SMRPs was low, PE-SMRPs can be a good indicator of the existence of MPM.

Abbreviations: AUC = area under the curve, CEA = carcinoembryonic antigen, CEIA = chemiluminescent enzyme immunoassay method, DOR = diagnostic odds ratio, ELISA = enzyme linked immunosorbent assay method, FN = false negative, FP = false positive, HGMB = high mobility group box protein 1, MM = malignant mesothelioma, MPM = malignant pleural mesothelioma, NLR = negative likelihood ratio, PE = Pleural effusions, PLR = positive likelihood ratio, QUADAS = Quality Assessment of Diagnostic Accuracy Studies, SMRPs = soluble mesothelin-related peptides, SROC = summary receiver operating characteristic, TN = true negative, TP = true positive.

Keywords: accuracy, diagnosis, mesothelin, mesothelioma, pleural effusion

1. Introduction

Malignant pleural mesothelioma (MPM) is a rapidly developing, almost fatal malignant tumor.[1] Although its occurrence is relatively rare, it is mainly associated with exposure to asbestos. Because the carcinogenicity of asbestos has been detected in recent years, it has been strictly prohibited in the United States, Europe, Australia, and so on.[2,3] However, in some industrialized countries, such as China, the production and use of asbestos are still not subject to strict control, which conferring the incidence of MPM will continue to increase in the coming decades.[3,4] A large study shows that the global MPM death burden is assumed to be about 38,400 a year, probably well above that.[5] However, due to the lack of a history of asbestos exposure and difficulty in obtaining pathological tissue for examination, the diagnosis of mesothelioma may be underestimated.[6] Exploring biomarkers of mesothelioma is helpful for screening and early diagnosis of mesothelioma and improving prognosis.

In diagnosis, the most widely studied biomarkers are mesothelin, osteopontin, fibulin-3, and high mobility group box protein 1.[5–10] Of which, soluble mesothelin-related peptides (SMRPs) is one of the most robust biomarkers at present for the diagnosis of MPM. Mesothelin is a 40-kD cell surface glycoprotein, which has putative functions in cell-to-cell adhesion.[11] It is highly overexpressed in cancers such as malignant mesothelioma, pancreatic or ovarian carcinoma, sarcomas, and in some gastrointestinal or pulmonary carcinoma.[12] Through 2 mechanisms of abnormal splicing events or an enzymatic cleavage from membrane-bound mesothelin, elevated amounts of SMRPs could be found in pleural effusions (PE).[13]
This provides the possibility for the application of PE-SMRPs in the diagnosis of MPM.

Actually, the diagnostic accuracy of SMRPs for MPM has been extensively studied, previous diagnostic experiments have shown that PE-SMRPs has higher specificity and good diagnostic efficacy. In particular, previous meta-analysis studies on SMRPs in blood and PE showed certain diagnostic value. Considering the small number of meta-analysis by Cui et al, we performed a meta-analysis to update the overall diagnostic accuracies of SMRPs in PE for identifying MPM.

2. Methods

Neither ethical approval nor consent to participate is applicable because this is a meta-analysis based on previous published studies.

2.1. Publication search strategy

The electric database included Medline, Embase, Web of Science, and Cochrane library, were systematically searched until August 24, 2018. The search keywords were as follow: “mesothelin” or “soluble mesothelin-related peptides” or “SMRPs” and “malignant pleural mesothelioma” or “mesothelioma” and “diagnosis” or “diagnostic value” or “diagnostic accuracy”. Studies were included when it provided diagnostic value of SMRPs in PE with both sensitivity and specificity for diagnosing MPM from other diseases included non-MPM malignancy, benign and undiagnosed disease.

2.2. Inclusion and exclusion criteria

Participants with PE were identified for inclusion. The diagnosis of MPM was proven on histopathology obtained by biopsy histopathology or pleural fluid cytology.

The exclusion criteria were: reviews, meta-analyses, case reports, letters, or expert opinions; SMRPs detected in other specimens, not in PE; studies published in languages other than English or Chinese; insufficient data to calculate sensitivity and specificity; duplicates; multiple publications.

2.3. Data extraction and risk of bias assessment

Two reviewers (R.G., Q.Y.) independently extracted data from included trials and assessed the risk of bias. The extracted data mainly included 9 aspects: the first author; the year of the paper published; the region the study performed; the study type; the measurement of SMRPs; the standard criteria of MPM; the number of all included patients; the number of PM patients; the number of true positive, false positive, false negative, and true negative. The risk of bias of the included studies was assessed by the Quality Assessment of Diagnostic Accuracy Studies (QUADAS) tool and rated as “low,” “unclear,” or “high” in several domains. The two reviewers search and evaluate the quality of the studies independently. Any disagreement was resolved by a third reviewer (F.W.).

2.4. Statistical analyses

The diagnostic efficacy for PE-SMRPs was pooled by the equation of sensitivity= true positive/ (true positive+ false negative), specificity= true negative/true negative + false positive). The diagnostic parameters were pooled by the random effect model according to the statistical heterogeneity across the included studies. The symmetric receiver operating characteristic curve (SROC) was drawn and area under the curve (AUC) was calculated through sensitivity vs. specificity for diagnosing MPM. P < .05 was considered statistically significant. Subgroup analysis was used to explore sources of heterogeneity. The subgroup analysis was carried out according to the risk of bias on patient selection, index test, reference standard, flow, and timing.

3. Results

This study was confirmed by the PROSPERO register (Registration ID CRD42018116506). We identified 349 studies, from which 123 were duplicates, showed in Figure 1. After screening titles and abstracts, 188 trials were excluded because of not meeting the inclusion criteria. Then, we obtained full-text articles for 38 citations, of which, 23 were excluded for reasons described in the Figure 1. Ultimately, 13 studies met our eligibility criteria were included in this meta-analysis, involving a total of 3359 patients. One study had 2 best diagnostic performances, which were considered 2 independent studies.

3.1. The characteristic of included studies

The included studies were published between 2007 and 2017. The average sample size was 240 (range from 92 to 1331) and the total of 3356 patients was included, of whom, 756 were MPM patients, 1061 were non-MPM malignant PE cases, and 1539 were benign PE cases. The characteristic of included studies was presented in Table 1.

Most studies used histopathologic immunohistochemistry method as diagnostic criteria for MPM, only 3 studies used PE cytology with biopsy histology as diagnostic criteria.

In term of the type of study, 8 studies were cohort study that enrolled a consecutive or random sample of patients, 4 studies were case-control designed, and 1 study was an unclear type that enrolled a consecutive sample of eligible patients with suspected disease and a portion of participants with known disease. In addition to 2 studies measuring SMRPs using the chemiluminescent enzyme immunoassay method, the rest of the studies used enzyme linked immunosorbent assay method. Finally, the cutoff values applied in 4 article were previously established.

3.2. Risk of bias

The assessment of risk of bias was presented in Figure 2, which indicated patient selection, and reference standard existed high risk on the bias. Five studies had a high risk of bias on patient selection because these studies were not a consecutive sample of patients enrolled. One study had an unclear risk of bias on patient selection because it was a retrospective study. Three studies had a high risk of bias on reference standard because these studies used pleural fluid cytology as diagnostic criteria, which remain a controversial subject according to guideline. Nine studies had an unclear risk of bias on index test because these studies selected the test threshold to optimize sensitivity for evaluating diagnostic effectiveness, however, this bias might be moderate risk for eligible sample sizes.
3.3. Diagnostic accuracy

We noted that sensitivity was 0.69 (95% confidence interval [CI]: 0.64 to 0.72), specificity was 0.90 (95% CI: 0.85–0.94), positive likelihood ratio (PLR) was 4.78 (95% CI: 3.52–6.50), negative likelihood ratio (NLR) was 0.30 (95% CI: 0.24–0.36), and diagnostic odds ratio (DOR) was 19.50 (95% CI: 12.14–31.33). A forest plot of sensitivity and specificity for 13 studies of PE SMRPs for the diagnosis of MPM was presented in Figure 3. I² values of sensitivity and specificity were 2.97 (P = .42), 89.88 (P < .001), respectively, indicating massive heterogeneity in specificity between studies.

![Flow chart of selection process in this study.](image)

3.4. Subgroup analysis

Preliminary analysis showed that the overall diagnostic ability of PE-SMRPs was not as good as predicted with high heterogeneity. According to quality assessment, patient selection, the applicability of diagnostic criteria, and presupposition of cutoff value might be the cause of high heterogeneity.

| Reference | Year | Country | Type           | Measurement  | Standard criteria | Total MPM | Cutoff (nmol/L) | TP | FP | FN | TN |
|-----------|------|---------|----------------|--------------|-------------------|-----------|-----------------|----|----|----|----|
| [17]      | 2012 | France  | Unclear        | ELISA        | Histology         | 101       | 61              | 24.05 | 39 | 0 | 22 | 40 |
| [18]      | 2013 | Italy   | Cohort         | ELISA        | Histology         | 104       | 34              | 19.6 | 20 | 2 | 14 | 68 |
| [19]      | 2013 | Italy   | Cohort         | ELISA        | Histology         | 275       | 52              | 9.3  | 39 | 27 | 13 | 196|
| [20]      | 2014 | Australia | Cohort       | ELISA        | Cytology or histology | 1331      | 183             | 20   | 123 | 57 | 60 | 1091|
| [21]      | 2007 | Australia | Case-control | ELISA        | Cytology or histology | 192       | 52              | 20   | 35 | 3 | 17 | 137|
| [22]      | 2009 | UK      | Cohort         | ELISA        | Histology         | 166       | 24              | 20   | 17 | 14 | 7 | 128|
| [23]      | 2013 | Italy   | Cohort         | ELISA        | Histology         | 177       | 57              | 12  | 42 | 17 | 15 | 103|
| [24]      | 2010 | Japan   | Case-control   | CLEIA        | Histology         | 96        | 23              | 8    | 16 | 23 | 7 | 50 |
| [25]      | 2013 | UK      | Cohort         | ELISA        | Histology         | 196       | 25              | 20   | 18 | 21 | 7 | 147|
| [26]      | 2017 | Japan   | Cohort         | CLEIA        | Cytology or histology | 240      | 32              | 20   | 18 | 28 | 14 | 180|
| [27]      | 2008 | US      | Case-control   | ELISA        | Histology         | 95        | 45              | 12.6 | 34 | 9 | 11 | 41 |
| [20]      | 2008 | France  | Case-control   | ELISA        | Histology         | 92        | 43              | 20.8 | 28 | 8 | 15 | 41 |
| [28]      | 2008 | France  | Case-control   | ELISA        | Histology         | 92        | 43              | 11.4 | 33 | 15 | 10 | 34 |

CLEIA = chemiluminescent enzyme immunoassay, ELISA = enzyme linked immunosorbent assay, FN = false negative, FP = false positive, MPM = malignant pleural mesothelioma, TN = true negative, TP = true positive.

*Unclear type: continuous or random inclusion of eligible patients with suspected disease and at the same time a portion of patients diagnosed with the disease.
For further discover the diagnostic value of PE-SMRPs, first of all, we divided the literature into 2 groups, one was a cohort study group and the other was the noncohort study group, and remove studies by PE cytological diagnosis. Unfortunately, the removal of 1 study of PE criteria from the noncohort group could not be analyzed by Stata for asymmetric data, so the 1 study was not removed. The diagnostic comparison of sensitivity, specificity, PLR, NLR, DOR, and AUC between cohort group and noncohort group were showed in Table 2. The forest plot and the SROC of the cohort group and noncohort group were showed Figure 4, respectively.

The PE-SMRPs in cohort group had better diagnostic value and lower heterogeneity compared with the noncohort group, the AUC of the SROC of the cohort group was 0.86 with pooled sensitivity of 0.68 (95% CI: 0.60–0.75) and the pooled specificity of 0.91 (95% CI: 0.87–0.94), which showed in Figure 3. However, there might be a medium specificity due to the absence of analysis of the influence of preset cutoff.

We divided the study into 2 groups according to whether the cutoff value was greater than or less than 15 nM. However, the results could not reduce the heterogeneity which indicated meaningless. The results were shown in the Table 3.

3.5. Publication bias

The Deek’s test showed P = .13, which revealed there was no publication bias in this meta-analysis, showed in Figure 5.

4. Discussion

Since the abundant application of asbestos in the 1960s, the incidence of MPM is increasing. For clinicians, MPM is still susceptible to miss due to its low morbidity and nonspecific clinical manifestations. Moreover, the most recommended approach to diagnose MPM necessitates invasive examination such as thoracoscopy, which is restricted to achieve for old patients and increased the difficulty for diagnosis. It was worth noting that carcinoembryonic antigen (CEA) is a very widely used tumor marker, and PE-CEA had a superior diagnostic value in the diagnosis of malignant PE, but its diagnostic value for MPM was low. So far, the diagnosis of MPM depends entirely on histopathological examination, without a single biological marker, which is unfavorable for early diagnosis. SMRPs is one of the most popular soluble biomarker among studies. Given fewer patients taking into account previous meta-analysis, we therefore, performed a meta-analysis to update the overall diagnostic accuracies of pleural SMRPs.

Our results suggest that PE-SMRPs had an unfavorable diagnostic performance with poor sensitivity (0.69 95% CI: 0.64, 0.72) and high specificity (0.90 95% CI: 0.85, 0.94). The AUC was 0.76 (95% CI: 0.72, 0.80) indicating that the overall accuracy was not as high as expected. On one hand, high specificity revealed that PE-SMRPs above the cutoff value could be helpful for confirming MPM, so further invasive examination, will be strongly supported. On the other hand, low sensitivity (0.69 95% CI: 0.64, 0.72) limited SMRPs’ diagnostic performance. When the level of SMRPs was lower than cutoff value, the diagnosis could not exclude from MPM. In general, the PLR is greater than 10 and the NLR is less than 0.1 to exclude the possibility of disease. This indicates that PE-SMRPs does not have the ability to diagnose MPM directly, but it also has certain diagnostic value. In view of the heterogeneity of specificity (I² = 89.88, P < .01), this reveal the existence of inclusion bias, so further subgroup analysis is requisite.

The QUADAS tool is the widest tool to assess the quality of included studies in systematic reviews of diagnostic accuracy studies for markedly heterogeneous results, which is improved to QUADAS-2. The results of QUADAS-2 (Fig. 1) suggested patient selection, the applicability of diagnostic criteria, and cutoff value might be the main cause of heterogeneity. Among them, the type of study and diagnostic criteria has a greater impact on the bias. The presence of an inappropriate exclusion of undiagnosed PE in the case–control study might result in overestimation of diagnostic accuracy. At the same time, diagnosis in cytology remains controversial because the sensitivity of cytology for the diagnosis of mesothelioma ranges from
30% to 75% and epithelioid mesothelial cells are easier to exfoliate tumor cells in the pleural cavity compared with sarcomatoid mesothelioma, which misleads the pathologist and reduces nonepithelioid mesothelioma diagnosis. To sum up, the cytology diagnosis is unreliable.

According to the type of study, 13 studies were divided into cohort study group and noncohort study group. It was worth noting that the study of cytological diagnosis was excluded in the cohort group but not in the noncohort group for calculating. The result showed cohort group provided better diagnostic accuracy than the noncohort group for MPM: AUC was 0.86 (95% CI: 0.83–0.89) compared with 0.76 (95% CI: 0.71–0.78), meanwhile, the heterogeneity of specificity in cohort group decline to 61.01 (P < .01). Therefore, we believe that the results of the cohort group can better reflect the actual diagnostic value of SMRPs.

However, the results of the cohort group also had low sensitivity, similar to previous study results, which might be related to its production. SMRPs is a glycoprotein on the surface of epithelial cells. The production of SMRPs is related to abnormal splicing events leading to the synthesis of a secreted protein or to an enzymatic cleavage from membrane-bound mesothelin, which could be found in PE. This characteristic might be the main reason for its low sensitivity of SMRPs for MPM. In the next place, SMRPs is most expressed in epithelial mesothelioma, while almost absent in sarcomatous type. When the proportion of sarcomatous or other types of mesothelioma is increased, combined with cytology diagnosis,
it leads to increased false negative rate and reduced sensitivity. Moreover, SMRPs content is less in the early stage of mesothelioma, so more cases of early (I II period) are included, which will also lead to poor sensitivity. Unfortunately, our study did not make a further assessment of the diagnostic effectiveness of SMRPs by stages.

As a meta-analysis, the cutoff values of the articles included in this study are quite different (from minimum 8nmol/L to

Figure 4. Forest plot of estimates of sensitivity and specificity and summary receiver operating characteristic curve for SMRPs in the cohort group and noncohort group. SMRPs = soluble mesothelin-related peptides.

| Table 2 | The comparison of diagnostic performance between cohort group and non-cohort group. |
|---|---|
| | Cohort | Noncohort |
| Studies (reference numbers) | [16–20,22,23,25] | [17,21,24,27,28] |
| Sensitivity (95% CI) | 0.68 (0.60,0.75) | 0.70 (0.63,0.76) |
| Heterogeneity (P value) | 22.6 ($P = .24$) | 0.00 ($P = .66$) |
| Specificity (95% CI) | 0.91 (0.87,0.94) | 0.89 (0.72,0.76) |
| Heterogeneity (P value) | 61.01 ($P < .01$) | 90.23 ($P < .01$) |
| PLR (95% CI) | 7.6 (5.3,10.9) | 6.5 (2.4,18.1) |
| NLR (95% CI) | 0.35 (0.28,0.43) | 0.34 (0.28,0.41) |
| DOR (95% CI) | 22 (14.32) | 19 (15.49) |
| AUC (95% CI) | 0.86 (0.83,0.89) | 0.67 (0.71,0.78) |

AUC = area under the curve, DOR = diagnostic odds ratio, NLR = negative likelihood ratio, PLR = positive likelihood ratio.

| Table 3 | The comparison of diagnostic performance between cutoff < 15nM and non-cutoff > 15nM. |
|---|---|
| | Cutoff < 15nM | Cutoff > 15nM |
| Studies (reference numbers) | [16–18,20–22,25,26,28] | [19,23,24,27,28] |
| Sensitivity (95% CI) | 0.75 (0.68,0.80) | 0.65 (0.60,0.69) |
| Heterogeneity (P value) | 0.00 ($P = .97$) | 0.00 ($P = .72$) |
| Specificity (95% CI) | 0.80 (0.72,0.87) | 0.94 (0.90,0.96) |
| Heterogeneity (P value) | 60.44 ($P < .01$) | 83.23 ($P < .01$) |
| PLR (95% CI) | 7.6 (5.3,10.9) | 6.5 (2.4,18.1) |
| NLR (95% CI) | 0.35 (0.28,0.43) | 0.34 (0.28,0.41) |
| DOR (95% CI) | 22 (14.32) | 19 (15.49) |
| AUC (95% CI) | 0.86 (0.83,0.89) | 0.67 (0.71,0.78) |

AUC = area under the curve, DOR = diagnostic odds ratio, NLR = negative likelihood ratio, PLR = positive likelihood ratio.
[maximum 24.1 nmol/L). Therefore, with 15 nmol/L as the boundary, we are divided included studies into 2 groups according to the cutoff values, which were presented in Table 3. The results showed that group (cutoff < 15 nmol/L) had higher diagnostic efficacy because it had higher AUC (0.76 (95% CI: 0.72, 0.80) vs. 0.67 (95% CI: 0.63, 0.71)), but lower specificity (0.80 [95% CI: 0.72, 0.87] vs. 0.94 [95% CI: 0.90, 0.96]) and lower PLR (3.8 [95% CI: 2.6, 5.6] vs. 10.3 [95% CI: 6.1, 17.3]), which meant that more suspicious MPM patients could be included in context of lower level of cutoff value of PE-SMRPS with a compromised ability to confirm MPM. In addition, we also compared the diagnostic effectiveness of actual experiments with different cutoff values, which showed in Table 4. The results revealed that AUC maintained from 0.767 to 0.878 with cutoff value (20 ± 0.4 nmol/L). Moreover, the specificity with cutoff value of 20 ± 0.4 nmol/L was generally higher than that of

| Range of cutoff | Studies reference | Cutoff (nmol/L) | Sensitivity (%) | Specificity (%) | AUC |
|----------------|------------------|----------------|----------------|----------------|-----|
| 20 ± 0.4 nmol/L | [18]             | 19.6           | 58.8           | 97.1           | 0.767∗ |
|                | [20]             | 20             | 58             | 96             | 0.815 (95% CI: 0.743, 0.874) |
|                | [22]             | 20             | 71             | 90             | 0.878 (95% CI: 0.807, 0.948) |
|                | [24]             | 20             | 72             | 87.5           | 0.85 (95% CI: 0.77, 0.94) |
|                | [25]             | 20.8           | 65.1           | 83.7           | 0.809 (95% CI: 0.718, 0.900) |
| 12 ± 0.6 nmol/L | [26]             | 11.4           | 76.7           | 69.4           | 0.809 (95% CI: 0.718, 0.900) |
|                | [27]             | 12.6           | 75.6           | 82             | 0.76 (95% CI: 0.662, 0.842) |
| 8 ± 0.6 nmol/L  | [19]             | 9.3            | 75.7           | 93.5           | 0.82 (95% CI: 0.748, 0.904) |
|                | [24]             | 8              | 70             | 68.4           | 0.75 (95% CI: 0.615, 0.884) |

∗No confidence intervals were provided in the original language.
AUC = area under the curve.

Figure 5. The Deek’s funnel plot of publication bias for SMRP in diagnosis of MPM. MPM = malignant pleural mesothelioma, SMRPs = soluble mesothelin-related peptides.
cutoff value of $12 \pm 0.6 \text{nmol/L}$ and $8 \pm 0.6 \text{nmol/L}$, with a range of $83.7\%$ to $97.1\%$. Altogether, we can conclude that MPM should be suspected when PE-SMRPs $>8 \text{nmol/L}$. When PE-SMRPs $>20 \text{nmol/L}$, MPM should be strongly suspected and further invasive procedures may be intended for patients.

There are several advantages to this study. First, we assessed the diagnostic performance of PE-SMRPs differentiating other diseases from MPM through subgroup analysis, which was helpful for application practice. Furthermore, our results were consistent with those of Cui et al, we involved a larger number of patients and the subgroup analysis of the type was carried out, which had more reference value.[7]

However, this study also presents several limitations. First of all, only full English studies included in the present study and exclusion of conference abstracts and non-English language studies might result in publication bias. However, the results of Deek's test showed that there was no publication bias in this study. Next, the subgroup analysis could not be carried out by meta-regression analysis, this was due to the cause that Stata could not be calculated for asymmetric data. Nonetheless, our results according to the consequence of QUADAS-2 showed that the direction of subgroup analysis was correct, which effectively reduced the heterogeneity and improved the diagnostic efficacy. Ultimately, SMRPs in PE can be a good indicator of the existence of MPM. Though poor sensitivity of SMRPs might miss sarcomatous or another type mesothelioma, high specificity of SMRPs could point out mesothelioma, which provides strong evidence for further invasive examination. It is necessary to further explore highly sensitive biomarkers combined with SMRPs for the diagnosis of MPM.

5. Conclusion
SMRPs in PE can be a good indicator of the existence of MPM, high level of pleural SMRPs are highly suspicious for MPM which provides supporting evidence for further invasive examinations to obtain biopsy results from the pleura. However, its sensitivity is too low to be used as an indicator of MPM screening.

Author contributions
Ruiyue Gao, Yue Qin collected the relevant papers and data. Ruiyue Gao performed statistical analysis and data synthesis. All authors contributed substantially to the study design, data interpretation, and the writing of the manuscript. Feng Wang gave critical comments and revised the paper. All authors approved the final version of the manuscript.

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