Large-scale remote fear conditioning: Demonstration of associations with anxiety using the FLARe smartphone app

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Abstract
Objectives: We aimed to examine differences in fear conditioning between anxious and nonanxious participants in a single large sample.

Materials and methods: We employed a remote fear conditioning task (FLARe) to collect data from participants from the Twins Early Development Study (n = 1,146; 41% anxious vs. 59% nonanxious). Differences between groups were estimated for their expectancy of an aversive outcome towards a reinforced conditional stimulus (CS+) and an unreinforced conditional stimulus (CS−) during acquisition and extinction phases.

Results: During acquisition, the anxious group (vs. nonanxious group) showed greater expectancy towards the CS−. During extinction, the anxious group (vs. nonanxious group) showed greater expectancy to both CSs. These comparisons yielded effect size estimates (d = 0.26–0.34) similar to those identified in previous meta-analyses.

Conclusion: The current study demonstrates that remote fear conditioning can be used to detect differences between groups of anxious and nonanxious individuals, which appear to be consistent with previous meta-analyses including in-person studies.

KEYWORDS
anxiety disorders, differential conditioning, extinction, remote study, smartphones

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0 INTRODUCTION

Fear conditioning models aversive associative learning, a key process involved in the development, maintenance, and treatment of anxiety disorders (Craske et al., 2018; Mineka & Oehlberg, 2008; Pittig et al., 2018). Differential fear conditioning tasks use neutral "conditional stimuli" (CS; e.g., images of shapes) and aversive "unconditional stimuli" (US; e.g., loud scream) to experimentally manipulate fear-based learning. During acquisition, one CS is reinforced (CS+) by repeatedly presenting it with the aversive US, while another nonreinforced CS is presented alone (CS−). After multiple presentations, the CS+ typically elicits a conditional response (e.g., sweating) reflecting anticipation of US onset, whereas the CS− does not. Conditional responses can include self-report, behavioral, and physiological/neurobiological changes (Lonsdorf et al., 2017). During extinction, the CS+ and CS− are repeatedly presented without the US. Extinction usually results in a decrease in conditional responses driven by the development of a competing association between the CS+ and safety (Bouton, 1993). Acquisition and extinction model the development and exposure-based treatment of anxiety, respectively.

Multiple studies have examined differences in fear conditioning between anxious and nonanxious participants, with varying patterns of response emerging across different anxiety disorders and outcome measures. During acquisition, anxious participants tend to show either greater responding to the CS+ and CS− (Blechert et al., 2007; Norrholm et al., 2011; Orr et al., 2000), or to the CS− only (Lissek et al., 2009, 2010; Rabinak et al., 2017). This inconsistency is reflected in two meta-analyses, as one found stronger responses to both the CS+ and CS− (Lissek et al., 2005) while the other, more recent analysis found stronger responses to the CS− only (Duits et al., 2015). These findings suggest poor inhibitory responding to safety (CS−) among anxious participants and potentially increased excitatory responding to threat (CS+). One explanation for poor inhibitory responding is that, during threatening or uncertain situations, anxious participants generalize their fear of the CS+ towards nonthreatening stimuli (i.e., the CS−; Duits et al., 2015, for a review of alternative explanations see Lissek et al., 2005). Findings are somewhat more consistent during extinction, with studies showing stronger responses to the CS+ for anxious participants compared with controls, suggesting that it is difficult for them to develop new inhibitory learning to a previously threatening cue (Blechert et al., 2007; Duits et al., 2015; Lissek et al., 2005; Michael et al., 2007; Norrholm et al., 2011).

Inconsistencies across studies are likely due to heterogeneity in the adopted methodology, sample, and analytical approach (Lonsdorf et al., 2017, 2019; Ney et al., 2018). In addition, effect size estimates from the most recent meta-analysis suggest differences between groups are modest (d = 0.3–0.35; Duits et al., 2015). Therefore, the typically small sample sizes used in fear conditioning research likely mean that many studies are underpowered to detect these effects. For example, no individual study included in Duits et al.’s (2015) meta-analysis had sufficient power to detect the effect sizes produced by combining samples. The largest study in the meta-analysis had 270 participants (Norrholm et al., 2013). Power calculations (using G*Power; Faul et al., 2007) show 352 participants would be required to detect similar effect sizes (d > 0.3, a = .05, 1 - β = .8) to those observed for the main outcomes in Duits et al. (2015).

To overcome issues around study heterogeneity and power, we developed a smartphone app, Fear Learning and Anxiety Response (FLARe), which delivers a fear conditioning task remotely via smartphone (Purves et al., 2019). Remote delivery removes many barriers to conducting large-scale experiments by vastly increasing the number of simultaneous assessments, and reducing the time and cost needed. FLARe assesses self-reported expectancy of the US (a loud scream) during CS presentations (geometric shapes) throughout acquisition and extinction. We have previously validated FLARe against standard in-person (laboratory) data collection, demonstrating within-person correlations of fear conditioning outcomes between laboratory and app delivery did not differ from those seen across time using the same delivery mode (Purves et al., 2019). FLARe presents a novel opportunity to examine differences between anxious and nonanxious individuals within a large sample without the confounds of task, sample, and analysis variability.

The current study is the first attempt to assess differences in fear conditioning between anxious versus nonanxious participants via data collected remotely from a single, large sample using a mobile app, FLARe. The primary analyses investigated differences in fear conditioning between anxious participants self-reporting current or lifetime, clinically relevant anxiety, and nonanxious participants reporting no such experience. In addition, secondary sensitivity analyses examined current and prior anxiety separately. To compare with previous research, we present mean discrimination (CS+ minus CS−) and CS-specific (CS+ or CS−) between-group differences in expectancy scores for each phase (acquisition and extinction). Based on subjective (i.e., self-reported) results from the most recent meta-analysis (Duits et al., 2015), we hypothesized: (i) during acquisition, anxious individuals (compared with nonanxious individuals) would show greater responses to the CS− and poorer discrimination scores, and (ii) during extinction, anxious individuals would display greater responses to both the CS+ and CS−.

2 METHODS

2.1 Participants

Participants were recruited from the Twins Early Development Study (TEDS) via email invitation. TEDS is a longitudinal birth cohort study of twins born in England and Wales between 1994 and 1996. The study initially recruited approximately 16,000 families and approximately 8,000 continue to participate (Rimfeld et al., 2019). The cohort continues to be roughly representative of the population in England and Wales with regard to ethnicity and family socio-economic factors.

To take part in the current study, participants needed an Android or iOS smartphone to download the FLARe app, which
delivered the fear conditioning task and collected detailed study information including informed consent. Ethical approval was granted by the King’s College London Psychiatry, Midwifery and Nursing Research Ethics Subcommittees (application PNM/09/10-104). All participants who completed the study received a £10 gift voucher as reimbursement.

Figure 1 illustrates sample size from recruitment to analyses. See supplementary information for anxiety and socioeconomic status comparisons between consenters and nonconsenters and Table S1 for information on twin relatedness. Participants were screened-out if they had a pre-existing heart condition, a neurological condition, an uncorrected hearing impairment, or were pregnant. Of the participants meeting screening criteria, 382 participants did not complete the task and a further 180 rated the US unpleasantness five out of 10 or lower (1=“No unpleasant at all”, 10=“Very unpleasant”). Participants were excluded from analyses (n = 738) for self-reporting they did not follow the task instructions (e.g., removing their headphones), or if the app detected it had (a) been closed or (b) average device volume was lower than 50%. The final number of participants who met criteria to be included in either group (anxious vs. nonanxious) was 1,146.

2.2 | Procedure

After downloading and logging into the app, participants were asked to complete consent and screening procedures. Eligible participants continued by supplying demographic information. Next, participants were given setup instructions (see Figure S1). They were instructed to complete the session alone, in a quiet room where they were unlikely to be disturbed. During the setup, the app detects whether (a) the mobile device is connected to headphones and (b) its volume is set to maximum. Participants are only able to begin the fear conditioning task once these two requirements are met. Task-specific instructions were provided following setup (see Figure S2).

We assessed self-reported expectancy of the US during each CS presentation (for average trial-by-trial expectancies see Figure S3). Stimulus choice for the CS (different sized circles) and US (a loud scream) was made because they could easily be delivered via smartphone and are often used in fear conditioning (Lonsdorf et al., 2017). During the first phase of the task, acquisition, participants completed a total of 24 pseudo-randomized trials (12 per CS). Each trial lasted 8 s (Figure 2b). Throughout each trial, one CS was presented superimposed on a context image of an outdoor scene (Context A; Figure 2a). After 2 s, expectancy ratings became available at the bottom of the screen. On 75% of CS+ trials, the US occurred during the final 500 ms of the trial. The US never occurred during CS− trials. Each trial was separated by an intertrial interval (ITI) where participants were instructed to focus on a fixation cross. ITI length was randomized to be 2, 2.5, or 3 s. After acquisition, participants had a 10-min break during which they completed the first set of questionnaires (not analyzed here, see Table S2 for details on all questionnaire measures collected during the study). Following the break, participants completed the second phase of the fear conditioning task, fear extinction. Extinction consisted of 36 trials in total (18 per CS). Trials followed the same format as acquisition, although the context image used was that of an indoor scene (Context B) and neither CS was paired with the US. Although not used here, FLARe has an additional optional phase delivered a day later to allow assessment of fear renewal in Context A (i.e., ABA conditioning). Renewal was not included to minimize participant burden. The AB context switch in the current study is comparable to studies included in Duits et al. (2015) that also switched context after acquisition as part of a return of fear paradigm (e.g., Milad et al., 2008, 2009, 2013). For each trial, measures of task compliance were collected via the app including headphone connection, volume and whether participants left the app. Following extinction, participants were redirected to a second set of questionnaires hosted on an external website (Qualtrics, Provo, UT) which contained all the self-report measures used in the current study. These questionnaires included further items about task compliance, one of which was used to determine whether participants removed their headphones.

2.3 | Measures

2.3.1 | Questionnaire measures

Two questionnaires were used to assign individuals to the anxious or nonanxious group.
Generalized Anxiety Disorder seven-item scale
Participants rated the frequency that they had experienced symptoms of anxiety over the past two weeks on a four-point scale ranging from "Not at all" (0) to "Nearly every day" (3). The Generalized Anxiety Disorder seven-item scale (GAD-7) has been shown to have good criterion validity for detecting anxiety disorders (Spitzer et al., 2006). Total scores range from 0 to 21, with scores of 10 or greater indicating clinically significant moderate to severe anxiety.

Self-reported lifetime diagnoses
Participants were asked "Have you ever been diagnosed with one or more of the following mental health problems by a professional, even

FIGURE 2 Visualization of experimental design implemented in the FLARe app. Schematic of overall task structure (a) with numbers representing the amount of times a stimulus is presented. Schematic of trial structure (b). CS: conditional stimuli. US: unconditional stimulus, a loud human scream played through headphones at a loud volume. Context: an outdoor scene (Context A) during the acquisition phase, an indoor scene (Context B) during the extinction phase.
if you don’t have it currently?” (for a full list of response options, see Table S3). Single-item measures assessing anxiety disorder diagnosis have been shown to have reasonable agreement (76.7%) with more detailed algorithm-based assessments (Davies et al., 2021).

For the primary analyses, participants were included in the anxious group if they had clinically significant current levels of anxiety (GAD-7 > 10; n = 299) or if they reported having a diagnosis of an anxiety disorder across the lifespan (n = 306). Lifetime anxiety disorder diagnosis included generalized anxiety disorder, social phobia, panic attacks, agoraphobia, specific phobia, obsessive compulsive disorder (OCD), or posttraumatic stress disorder (PTSD). Diagnosis of OCD and PTSD were included to reflect the Duits et al. (2015) meta-analysis which was based on the DSM-IV (American Psychiatric Association, 1994), though no participants ended up reporting a diagnosis of OCD or PTSD (see Table S4 for a breakdown of anxiety diagnoses). Participants reporting that a clinician had diagnosed them as suffering from panic attacks (n = 22) were also included in the anxious group. Panic attacks are not, however, an official anxiety disorder diagnosis but rather a key symptom for some anxiety disorders, in particular panic disorder. This terminology was chosen to represent how mental health disorders are commonly referred to rather than their strict DSM classification (i.e., panic disorder) to help participants recognize and self-report their diagnosis. Single-item measures of self-reported, clinician diagnoses of panic attacks have been shown to have moderate agreement with algorithm-based measures of panic disorder (65.4%; Davies et al., 2021).

The two groups identified through the GAD-7 or by reporting lifetime diagnoses overlapped considerably (n = 132), resulting in the total number of participants in the anxious group being 473. Both groups were included to maximize power, but sensitivity analyses were also conducted using participants with current (n = 299) and prior (i.e., self-reported lifetime diagnosis, but GAD-7 < 10; n = 174) anxiety separately. Participants were included in the nonanxious group if their GAD-7 scores were below five (cut-off for mild anxiety) and did not report a lifetime mental health diagnosis.

2.3.2 | Fear conditioning measures

Expectancy ratings
During each trial, participants were asked to rate their certainty that the trial would end with the occurrence of a loud scream (US). This “expectancy rating” was made using a nine-point scale ranging from one (“certain no scream”) to nine (“certain scream”), with five indicating uncertainty (“uncertain”). Expectancy ratings are a valid index of fear conditioning (Boddez et al., 2013) commonly used in investigations comparing anxious and nonanxious individuals (Blechert et al., 2007; Lissek et al., 2009; Norrholm et al., 2011). For both phases, mean expectancy ratings for each stimulus (CS+/CS−) were calculated to index participants’ conditioning and extinction. In addition, differences between the mean expectancy ratings for each stimulus (CS+ minus CS−) were calculated to index discrimination learning for both phases.

2.4 | Statistical analyses
An analysis of variance (ANOVA) was conducted twice for both phases. The first ANOVA tested mean expectancy ratings for the CS using a two-factor mixed-design with group (anxious vs. nonanxious) and stimulus (CS+ vs. CS−) entered as between-subjects and within-subjects factors, respectively. The second tested CS-discrimination scores using a one-way between-subjects design where group (anxious vs. nonanxious) was entered as the between-subjects factor. For each ANOVA, follow-up tests were conducted for pairwise comparisons using Tukey’s Honestly Significant Difference (HSD) test to control for multiple comparisons. Cohen’s d effect size estimates were calculated to indicate standardized differences between means allowing for comparisons with Duits et al.’s (2015) meta-analysis. This process was conducted for the primary analysis comparing participants with current and/or lifetime anxiety to nonanxious participants. Sensitivity analyses were then run, considering current anxiety and prior anxiety cases only. All analyses were conducted in R (3.6.1) using the Stats (3.2.1) and Psych (1.8.12) packages.

Two additional sensitivity analyses were conducted (see supplementary information Tables S5–S7 and Figure S4). The first considered only one participant from complete pairs of twins to see whether the clustered nature of the data was impacting results. The second analyses included excluded participants to assess whether their data introduced considerable noise, or whether it appears to add further sensible/usable information. This was done due to the fact that the majority of excluded participants had been removed for self-reporting headphone removal, yet the point at which they removed their headphones could not be determined (e.g., during extinction or after many acquisition trials).

3 | RESULTS
Females were three times as likely as males to be in the anxious group (odds ratio = 2.96; 95% confidence interval = 2.21–3.98, p < .001; Table 1). The two groups were of similar age, with a mean difference in age of approximately 1.5 months (d = 0.15; t(998.34) = 2.46, p = .01).

On average, participants found the US highly unpleasant (M = 8.97, SE = 0.03). Mean GAD-7 scores were significantly higher for excluded participants (n = 738, M = 6.31, SE = 0.20) compared to participants that were not excluded (n = 1625, M = 5.43, SE = 0.13; d = 0.17; t(1319.4) = 3.69, p < .001).

3.1 | Acquisition
For the primary analyses, there was a statistically significant group × stimulus type interaction on mean expectancy ratings (F(1, 2228) = 26.06, p < .001; Table 2). Post hoc tests showed that, compared with the nonanxious group, the anxious group showed (a) no
significant difference in expectancy ratings towards the CS+ \((d = -0.14, p = .086)\) and (b) significantly higher mean expectancy ratings towards the CS− \((d = 0.28, p < .001;\) see Table 3 for mean expectancy scores). There was a statistically significant main effect of group on expectancy discrimination scores indicating that, compared with the non-anxious group, the anxious group had lower expectancy discrimination scores \((d = -0.25; F(1, 1144) = 17.20, p < .001).\) Figure 3 illustrates effect sizes for both phases of the fear conditioning task, and also indicates the findings for subjective outcome measures from the latest meta-analysis for comparison (Duits et al., 2015). Both secondary sensitivity analyses, that is, looking at participants with current and prior anxiety separately, followed the same pattern of results as the primary analyses.

### 3.2 | Extinction

There was no statistically significant group × stimulus type interaction on mean expectancy ratings \((F(1, 2288) = 2.25, p = .13;\) Table 2). There were, however, significant main effects of group \((F(1, 2288) = 51.34, p < .001),\) and stimulus \((F(1, 2288) = 293.74, p < .001).\) Post hoc Tukey tests showed that compared with the nonanxious group, the anxious group had higher mean expectancy ratings towards conditional stimuli \((d = 0.29, p < .001).\) Separately, the anxious group had significantly larger expectancy ratings for both the CS+ \((d = 0.34, p < .001)\) and CS− \((d = 0.26, p < .001).\) There was a statistically significant main effect of group on expectancy discrimination scores indicating that the anxious group had significantly higher expectancy discrimination scores than the nonanxious group \((d = 0.14; F(1, 1144) = 5.56, p = .019).\) Both secondary sensitivity analyses followed the same pattern of results as the primary analyses. However, discrimination scores for participants with prior anxiety only were not significantly different from the nonanxious group \((d = 0.11; F(1, 845) = 1.68, p = .195).\)

Two additional sensitivity analyses assessing the impact of excluding twin pairs and including participants who disregarded instructions (see supplementary information Tables S5–S7 and Figure S4) showed that the pattern of effects remained the same, though effect sizes varied slightly.

### 4 | DISCUSSION

Using a novel remote fear conditioning task, we examined differences in expectancy ratings during acquisition and extinction between anxious and nonanxious individuals in a large sample of young adults. During acquisition, anxious individuals had larger expectancy ratings towards the CS− \((d = 0.29)\) and smaller discrimination scores \((d = -0.25)\) compared with nonanxious individuals. During extinction, anxious individuals had larger expectancy ratings towards the CS+ \((d = 0.34)\) and CS− \((d = 0.26)\) compared with nonanxious individuals. In line with our hypotheses, our findings followed the same pattern of effects seen for subjective ratings in the most recent meta-analysis of fear conditioning in the anxiety disorders. Secondary sensitivity analyses showed these effects still stood when analyzing participants with current and prior anxiety separately. Significant anxiety-related differences were also found for discrimination scores during extinction. However, the effect size for this difference was very small \((d = -0.14)\) and was not observed in the prior anxiety only sensitivity analyses.

During acquisition, no anxiety-related difference in expectancy ratings was observed for the CS+. This could have been due to a high reinforcement rate (75%) causing a "strong situation" (Lissek et al., 2006) whereby ambiguity concerning the likelihood of the US occurring was low. In such a case, it is possible that most participants, regardless of anxiety status, would give high expectancy ratings. This may be further explained by a ceiling effect where our expectancy rating scale did not allow for enough variation in confidence levels regarding the upcoming US. As such, both anxious and nonanxious

### Table 1 Sample characteristics for each group indicating: The number and proportion of males and females; means and standard errors for age and GAD-7 scores; the number of participants meeting different criteria for inclusion in the anxious group

| Group | Sample characteristics | Number meeting anxiety group criteria |
|-------|------------------------|---------------------------------------|
|       | Sex (%) | Group means (SE) |                             |                          |                          |                          |                          |
|       | Females | Males | Age | GAD-7 | Number meeting anxiety group criteria |               |                          |                          |                          |                          |
|       |         |       |     |       | Self-reported lifetime diagnosis | GAD-7 diagnosis | GAD-7 diagnosis & GAD-7 diagnosis |
| Current and/or lifetime anxiety | 473 | 390 (0.82) | 83 (0.18) | 23.58 (0.04) | 10.77 (0.24) | 174 | 167 | 132 |
| Current anxiety | 299 | 251 (0.84) | 48 (0.16) | 23.6 (0.05) | 14.09 (0.18) | 0 | 167 | 132 |
| Prior anxiety only | 174 | 139 (0.8) | 35 (0.2) | 25.53 (0.06) | 5.06 (0.2) | 174 | 0 | 0 |
| Nonanxious | 673 | 413 (0.61) | 260 (0.39) | 23.7 (0.03) | 1.61 (0.05) | 0 | 0 | 0 |

Abbreviations: GAD-7, Generalized Anxiety Disorder seven-item scale; SE, standard error.
participants may have been constrained to one or two rating levels towards the "certain" end of the scale. The largest difference in expectancy ratings was seen towards the CS−, with anxious participants displaying greater US expectancy to the safety stimulus than nonanxious participants. This finding suggests anxious individuals have a tendency to generalize their threat response from threatening stimuli (CS+) to nonthreatening but perceptually similar stimuli (CS−). Further evidence of this "overgeneralization" is provided through anxious participants' poor discriminatory learning. Overgeneralization is thought to maintain/exacerbate anxiety symptoms by increasing the number of threatening cues an anxious individual perceives in their environment (Lissek, 2012).

During extinction, the anxious group showed greater expectancy ratings for both CSs compared with the nonanxious group. This provides evidence that anxious individuals are resistant to the extinction of both conditioned (CS+) and generalized (CS−) fear. Difficulty extinguishing generalized fear poses a challenge for individuals undergoing exposure therapy. For example, for a patient with a phobia of dogs, exposure therapy focused on a single dog may help reduce their anxiety towards that breed but leave the patient with a generalized phobia towards other dogs. Previous research has shown that strengthening inhibitory learning (i.e., learning that a stimulus can be safe) towards a variety of related stimuli can improve extinction in conditioning tasks and exposure outcomes for anxious individuals (Carpenter et al., 2019; Craske et al., 2008). In practice, a clinician might decide to treat a patient with a phobia of dogs by eventually exposing them to a number of different breeds, colors and sizes of dog to reduce the patient’s symptoms.

Our findings were consistent with a previous meta-analysis (Duits et al., 2015), highlighting important differences in fear conditioning processes between anxious and nonanxious individuals. Though effect sizes were modest, these differences were observed in participants reporting both current and prior anxiety and provide further evidence of fear conditioning’s ability to model differences between healthy and at-risk individuals using expectancy rating data (Boddez et al., 2013). However, the ability to use individual differences in fear conditioning response to predict differences in anxiety or treatment response (i.e., predictive validity) remains the best test of whether findings from human fear conditioning research will help us understand the development and treatment of anxiety (Carpenter et al., 2019). Studies have used overgeneralization of fear and deficits in extinction learning to assess risk (Lommen et al., 2013; Sijbrandij et al., 2013), or predict treatment outcomes (Forcadell et al., 2017; Raedt et al., 2020; Waters & Pine, 2016), in anxiety disorder patients with varying levels of success. Inconsistent findings could relate to the relatively small effect sizes seen for the differences between groups, as demonstrated by effect sizes in our study and a previous meta-analysis.

| TABLE 2 | For both phases (acquisition/extinction), results for two ANOVAs testing (i) mean expectancy ratings with group (anxious/nonanxious) as a between-subjects factor and stimulus (CS+/CS−) as a within-subjects factor; (ii) CS-discrimination scores where group (anxious/nonanxious) was entered as the between-subjects factor |
|---|---|---|---|---|---|---|---|---|---|---|---|---|
| Whole phase means | Extinction | Discrimination |
| | df | F | p | df | F | p | df | F | p | df | F | p |
| **Current and/or lifetime anxiety** | | | | | | | | | | | | |
| Intercept | 1 | 27.630 | <.001 | 1 | 8318.8 | <.001 | 1 | 3241.6 | <.001 | 1 | 726.79 | <.001 |
| Group | 1 | 3.13 | .077 | 1 | 51.34 | <.001 | 1 | 17.20 | <.001 | 1 | 5.56 | .019 |
| Stimulus | 1 | 4910.53 | <.001 | 1 | 293.74 | <.001 | | | | | | |
| Group × stimulus | 1 | 26.06 | <.001 | 1 | 2.25 | .134 | | | | | | |
| **Current anxiety** | | | | | | | | | | | | |
| Intercept | 1 | 23.380.02 | <.001 | 1 | 6966.08 | <.001 | 1 | 2814.1 | <.001 | 1 | 606.57 | <.001 |
| Group | 1 | 2.07 | .151 | 1 | 43.16 | <.001 | 1 | 16.5 | <.001 | 1 | 5.53 | .019 |
| Stimulus | 1 | 4257.78 | <.001 | 1 | 247.86 | <.001 | | | | | | |
| Group × stimulus | 1 | 24.96 | <.001 | 1 | 2.26 | .133 | | | | | | |
| **Prior anxiety only** | | | | | | | | | | | | |
| Intercept | 1 | 22.454.83 | <.001 | 1 | 6677.96 | <.001 | 1 | 2788.12 | <.001 | 1 | 543.65 | <.001 |
| Group | 1 | 2.11 | .147 | 1 | 24.61 | <.001 | 1 | 5.34 | .019 | 1 | 1.68 | .195 |
| Stimulus | 1 | 4331.07 | <.001 | 1 | 230.81 | <.001 | | | | | | |
| Group × stimulus | 1 | 8.6 | .003 | 1 | 0.71 | .398 | | | | | | |

Note: Degrees of freedom, F statistics and p values for each ANOVA at acquisition and extinction. Significant p values (p < .05) are emphasized in bold. Abbreviations: ANOVA, analysis of variance; CS, conditional stimulus.
TABLE 3 Mean expectancy rating scores for each group and each phase (acquisition/extinction) with Tukey HSD estimates, Cohen’s $d$ estimates, and $p$ values for the difference between groups for all fear conditioning outcomes

| Group                          | Acquisition CS+ | Acquisition CS− | Discrimination | Extinction CS+ | Extinction CS− | Discrimination |
|-------------------------------|-----------------|-----------------|----------------|----------------|----------------|----------------|
| Current and/or lifetime anxiety |                 |                 |                |                |                |                |
| Group means (SE) Anxious (n = 473) | 6.95 (0.07)     | 3.12 (0.07)     | 3.83 (0.12)    | 3.49 (0.08)    | 2.38 (0.07)    | 1.11 (0.06)    |
| Nonanxious (n = 673)          | 7.15 (0.05)     | 2.7 (0.05)      | 4.45 (0.09)    | 2.97 (0.05)    | 2.04 (0.04)    | 0.93 (0.05)    |
| HSD                           | −0.2            | 0.42            | −0.62          | 0.52           | 0.34           | 0.18           |
| (95% CIs)                     | (−0.42 to 0.02) | (0.2−0.64)      | (−0.91 to −0.33) | (0.3−0.74)     | (0.12−0.56)    | (0.03−0.33)    |
| Cohen’s $d$                   | −0.14           | 0.29            | −0.25          | 0.34           | 0.26           | 0.14           |
| (95% CIs)                     | (−0.26 to −0.03) | (0.17−0.4)     | (−0.37 to −0.13) | (0.22−0.46)    | (0.14−0.38)    | (0.02−0.26)    |
| Adjusted $p$ value            | .086            | <.001           | <.001          | <.001          | <.001          | <.019          |
| Current anxiety               |                 |                 |                |                |                |                |
| Group means (SE) Anxious (n = 299) | 6.9 (0.09)      | 3.15 (0.09)     | 3.74 (0.16)    | 3.53 (0.1)     | 2.39 (0.09)    | 1.14 (0.08)    |
| Nonanxious (n = 673)          | 7.15 (0.05)     | 2.7 (0.05)      | 4.45 (0.09)    | 2.97 (0.05)    | 2.04 (0.04)    | 0.93 (0.05)    |
| HSD                           | −0.25           | 0.45            | −0.7           | 0.55           | 0.35           | 0.21           |
| (95% CIs)                     | (−0.5 to 0)     | (0.2−0.71)      | (−1.04 to −0.36) | (0.3−0.8)     | (0.1−0.6)      | (0.03−0.38)    |
| Cohen’s $d$                   | −0.18           | 0.32            | −0.28          | 0.37           | 0.27           | 0.16           |
| (95% CIs)                     | (−0.31 to −0.04) | (0.18−0.45)    | (−0.42 to −0.14) | (0.23−0.5)    | (0.14−0.41)    | (0.03−0.3)     |
| Adjusted $p$ value            | .058            | <.001           | <.001          | <.001          | <.001          | <.001          |
| Prior anxiety only            |                 |                 |                |                |                |                |
| Group means (SE) Anxious (n = 174) | 7.03 (0.1)      | 3.06 (0.12)     | 3.97 (0.19)    | 3.43 (0.12)    | 2.36 (0.11)    | 1.07 (0.1)     |
| Nonanxious (n = 673)          | 7.15 (0.05)     | 2.7 (0.05)      | 4.45 (0.09)    | 2.97 (0.05)    | 2.04 (0.04)    | 0.93 (0.05)    |
| HSD                           | −0.12           | 0.36            | −0.48          | 0.45           | 0.32           | 0.13           |
| (95% CIs)                     | (−0.42 to 0.18) | (0.06−0.66)     | (−0.88 to −0.08) | (0.17−0.74)   | (0.04−0.61)    | (−0.07 to 0.33) |
| Cohen’s $d$                   | −0.09           | 0.26            | −0.2           | 0.32           | 0.27           | 0.11           |
| (95% CIs)                     | (−0.26 to 0.08) | (0.09−0.42)     | (−0.37 to −0.03) | (0.16−0.49)   | (0.1−0.44)     | (−0.06 to 0.28) |
| Adjusted $p$ value            | .722            | .011            | .019           | <.001          | <.019          | .195           |

Note: Significant $p$ values ($p < .05$) are emphasized in bold.
Abbreviations: CIs, confidence intervals; CS, conditional stimulus; HSD, Tukey’s honestly significant difference; SE, standard error.

(Duits et al., 2015), which suggests a substantial proportion of variance is unexplained. Larger sample sizes afforded by remote research may improve our detection and prediction of individual differences in anxiety status and treatment response.

Using FLARe, we quickly collected data from a large sample of participants. This allowed us to conduct what we believe to be the largest human fear conditioning study to date. FLARe can easily be adopted in a variety of contexts, including clinical settings. Future research should take advantage of these benefits to reach clinical samples, study differences between specific types of anxiety disorders, and explore unique and interacting contributions between fear conditioning outcomes and other key processes underlying anxiety. In addition, the power afforded by large samples allows the use of more complex research methods, such as longitudinal studies including cohorts, treatment trials, and genetically sensitive designs.

4.1 | Limitations

Though remote research vastly increases ease of data collection, control over participant behavior is diminished. Many participants were excluded for not following task instructions (reported in a post-experiment survey), primarily headphone removal during testing. Self-reported anxiety scores indicated this set of participants were more anxious than participants who followed instructions; excluding them may have impacted our effect sizes.

While our findings replicated those from a meta-analysis (Duits et al., 2015), key differences in methodology should be highlighted. First, sample characteristics, such as age, sex, and distribution of anxiety diagnoses, differ across the two studies, as does the method used to obtain them. Previous meta-analyses included studies using clinician assessment, where as we used self-report measures to group participants. Like clinical interviews, self-report measures have
limitations relating to concerns around accuracy and memory of reporting. However, the GAD-7 has been shown to have good face validity for identifying individuals with anxiety disorders (Spitzer et al., 2006) and self-reported anxiety diagnoses have been shown to have reasonable agreement (76.7%) with an algorithm-based measure of anxiety disorders (Davies et al., 2021). Agreement for self-reported anxiety diagnoses was lower when looking at specific anxiety disorders separately, which was avoided in the current article. The effects of anxiety on patterns of fear conditioning were replicated in our study, suggesting that they generalize to groups selected using self-report measures. The meta-analysis also focused specifically on participants with current anxiety, whereas our study looked at participants with current and/or prior anxiety. Secondary sensitivity analyses in our sample showed similar patterns of results when current and prior anxiety were looked at separately. However, smaller effect sizes were observed when analyses were restricted to

FIGURE 3 Barplots of the effect sizes (d) per stimulus type, reflecting the standardized mean difference in expectancy ratings between anxious participants minus nonanxious participants during acquisition and extinction. Error bars display the unadjusted 95% confidence interval of the effect size estimate. Stars reflect the significance level of the difference between groups calculated using Tukey HSD which adjusts for multiple comparisons; *p < .05, **p < .01, ***p < .001. Red diamonds indicate the effect size estimates (d) for subjective ratings from the Duits et al. (2015) meta-analysis. CS, conditional stimulus; HSD, honestly significant difference.
those with prior anxiety only, suggesting anxiety-related differences in fear conditioning are particularly impacted by current anxiety.

Second, our findings were compared to subjective ratings from the previous meta-analysis, as expectancy ratings fall under this classification. The meta-analysis, however, included other subjective measures such as affective ratings. The current study used expectancy ratings due to their validity as an index of fear conditioning (Boddez et al., 2013) and the ease with which they could be collected remotely. Different outcome measures are considered to reflect different dimensions of fear/threat responses (Constantinou et al., 2020), making it beneficial to employ multiple outcome measures, a factor which should be considered in future studies.

Finally, our experimental design (e.g., contexts, stimuli, and trial lengths) could have impacted our results. For example, some evidence suggests that switching contexts from acquisition to extinction (AB conditioning) may attenuate expectancy ratings during early extinction (Effting & Kindt, 2007). In addition, there may have been an effect of measurement order. Anxiety disorder questionnaires were collected post-extinction and may have been impacted by state arousal. Our findings, therefore, may not extend to studies employing different designs.

5 | CONCLUSIONS

Our study used a remote fear conditioning task to examine differences between anxious and nonanxious individuals in a single, large sample. Results replicated findings from a previous meta-analysis, despite methodological differences. This consistency of findings, from meta-analysis to single study, improves our confidence in fear conditioning’s ability to differentiate healthy and anxious individuals. FLARe offers an exciting opportunity to enable studies in larger and harder to reach samples through remote fear conditioning and, we hope, will become a useful tool for future research.

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CONFLICT OF INTERESTS

The authors declare that there are no conflict of interests.

PEER REVIEW

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DATA AVAILABILITY STATEMENT

The data has not been made available on a permanent third-party archive; for information on how to request access to the data, see http://www.teds.ac.uk/researchers/teds-data-access-policy.

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REFERENCES

American Psychiatric Association. (1994). Diagnostic and statistical manual of mental disorders. 4th ed.
Blechert, J., Michael, T., Friends, N., Margraf, J., & Wilhelm, F. H. (2007). Fear conditioning in posttraumatic stress disorder: Evidence for delayed extinction of autonomic, experiential, and behavioural responses. Behaviour Research and Therapy, 45(9), 2019–2033. https://doi.org/10.1016/j.brat.2007.02.012
Boddez, Y., Baeyens, F., Luyten, L., Vansteenevengen, D., Hermans, D., & Beckers, T. (2013). Rating data are underrated: Validity of US expectancy in human fear conditioning. Journal of Behavior Therapy and Experimental Psychiatry, 44(2), 201–206. https://doi.org/10.1016/j.jbtep.2012.08.003
Bouton, M. E. (1993). Context, time, and memory retrieval in the interference paradigms of Pavlovian learning. Psychological Bulletin, 114(1), 80–99. https://doi.org/10.1037/0033-2909.114.1.80
Carpenter, J. K., Pinaire, M., & Hofmann, S. G. (2019). From extinction learning to anxiety treatment: Mind the gap. Brain Sciences, 9(7), 164. https://doi.org/10.3390/brainsci9070164
Constantinou, E., Purves, K. L., McGregor, T., Lester, K. J., Barry, T. J., Treanor, M., & Eley, T. C. (2020). Measuring fear: Association among different measures of fear learning. Journal of Behavior Therapy and Experimental Psychiatry, 101618. https://doi.org/10.1016/j.jbtep.2020.101618
Craske, M. G., Hermans, D., & Vervliet, B. (2018). State-of-the-art and future directions for extinction as a transnational model for fear and anxiety. Philosophical Transactions of the Royal Society of London, Series B: Biological Sciences, 373(1742), 1–10. https://doi.org/10.1098/rstb.2017.0025
Craske, M. G., Kircanski, K., Zelikowsky, M., Mystkowski, J., Chowdhury, N., & Baker, A. (2008). Optimizing inhibitory learning during exposure therapy. Behaviour Research and Therapy, 46(1), 5–27. https://doi.org/10.1016/j.brat.2007.10.003
Davies, M. R., Buckman, E. J. J., Adey, B. N., Armour, C., Bradley, J. R., Curzons, S. C. B., & Eley, T. C. (2021). Comparison of algorithm-based versus single-item phenotyping measures of depression and anxiety disorders in the GLAD Study cohort. MedRxiv. https://doi.org/10.1101/2021.01.08.21249434
Duijs, P., Cath, D. C., Lissek, S., Hox, J. J., Hamm, A. O., Engelhard, I. M., van den Hout, M. A., & Baas, J. M. P. (2015). Updated meta-analysis
of classical fear conditioning in the anxiety disorders. Depression and Anxiety, 32(4), 239–253. https://doi.org/10.1002/da.22353

Effting, M., & Kindt, M. (2007). Contextual control of human fear associations in a renewal paradigm. Behaviour Research and Therapy, 45(9), 2002–2018. https://doi.org/10.1016/j.brat.2007.02.011

Faul, F., Erdfelder, E., Lang, A.-G., & Buchner, A. (2007). G*Power 3: A flexible statistical power analysis program for the social, behavioral, and biomedical sciences. Behavior Research Methods, 39(2), 175–191. https://doi.org/10.3758/BF03193146

Forcadell, E., Torrents-Rodas, D., Vervliet, B., Leiva, D., Tortella-Feliu, M., & Fullana, M. A. (2017). Does fear extinction in the laboratory predict outcomes of exposure therapy? A treatment analogy study. International Journal of Psychophysiology, 121, 63–71. https://doi.org/10.1016/j.ijpsycho.2017.09.001

Lissek, S. (2012). Toward an account of clinical anxiety predicated on basic, neurally mapped mechanisms of Pavlovian fear-learning: The case for conditioned overgeneralization. Depression and Anxiety, 29(4), 257–263. https://doi.org/10.1002/da.21922

Lissek, S., Pine, D. S., & Grillon, C. (2006). The strong situation: A potential impediment to studying the psychobiology and pharmacology of anxiety disorders. Biological Psychology, 72(3), 265–270. https://doi.org/10.1016/j.biopsycho.2005.11.004

Lissek, S., Powers, A. S., McClure, E. B., Phelps, E. A., Woldehawariat, G., Grillon, C., & Pine, D. S. (2005). Classical fear conditioning in the anxiety disorders: a meta-analysis. Behaviour Research and Therapy, 43(11), 1391–1424. https://doi.org/10.1016/j.brat.2004.10.007

Lissek, S., Rabin, S., Heller, R. E., Lukenbaugh, D., Geraci, M., Pine, D. S., & Grillon, C. (2010). Overgeneralization of conditioned fear as a pathogenic marker of panic disorder. The American Journal of Psychiatry, 167(1), 47–55. https://doi.org/10.1176/appi.ajp.2009.09030410

Lissek, S., Rabin, S. J., McDowell, D. J., Dvir, S., Bradford, D. E., Geraci, M., Pine, D. S., & Grillon, C. (2009). Impaired discriminative fear-conditioning resulting from elevated fear responding to learned safety cues among individuals with panic disorder. Behaviour Research and Therapy, 47(2), 111–118. https://doi.org/10.1016/j.brat.2008.10.017

Lommen, M. J. J., Engelhard, I. M., Sijbrandij, M., van den Hout, M. A., & Hermans, D. (2013). Pre-trauma individual differences in extinction learning predict posttraumatic stress. Behaviour Research and Therapy, 51(2), 63–67. https://doi.org/10.1016/j.brat.2012.11.004

Lonsdorf, T. B., Klingelhofer-Jens, M., Andreattia, M., Beckers, T., Chalkia, A., Gerlicher, A., & Merz, C. J. (2019). Navigating the garden of forking paths for data exclusions in fear conditioning research. eLife. https://doi.org/10.7554/eLife.52465

Lonsdorf, T. B., Menz, M. M., Andreattia, M., Fullana, M. A., Golkar, A., Haaker, J., Heitland, I., Hermann, A., Kuhn, M., Kruse, O., Meir Drexl, S., Meulders, A., Nees, F., Pittig, A., Richter, J., Römer, S., Shibian, Y., Schmitz, A., Straube, B., … Merz, C. J. (2017). Don’t fear “fear conditioning”: Methodological considerations for the design and analysis of studies on human fear acquisition, extinction, and return of fear. Neuroscience and Biobehavioral Reviews, 77, 247–285. https://doi.org/10.1016/j.neubiorev.2017.02.026

Michael, T., Blechert, J., Friends, N., Magruf, J., & Wilhelm, F. H. (2007). Fear conditioning in panic disorder: Enhanced resistance to extinction. Journal of Abnormal Psychology, 116(3), 612–617. https://doi.org/10.1037/0021-843X.116.3.612

Milad, M. R., Furtak, S. C., Greenberg, J. L., Keshaviah, A., Im, J. J., Falkenstein, M. J., Jenike, M., Rauch, S. L., & Wilhelm, S. (2013). Deficits in conditioned fear extinction in obsessive-compulsive disorder and neurobiological changes in the fear circuit. JAMA Psychiatry, 70(6), 608–618. https://doi.org/10.1001/jamapsychiatry.2013.914

Milad, M. R., Orr, S. P., Lasko, N. B., Chang, Y., Rauch, S. L., & Pitman, R. K. (2008). Presence and acquired origin of reduced recall for fear extinction in PTSD: results of a twin study. Journal of Psychiatric Research, 42(7), 515–520. https://doi.org/10.1016/j.jpsychires.2008.01.017

Milad, M. R., Pitman, R. K., Ellis, C. B., Gold, A. L., Shin, L. M., Lasko, N. B., Zeidan, M. A., Handwerger, K., Orr, S. P., & Rauch, S. L. (2009). Neurobiological basis of failure to recall extinction memory in posttraumatic stress disorder. Biological Psychiatry, 66(12), 1075–1082. https://doi.org/10.1016/j.biopsych.2009.06.026

Mineka, S., & Oehlberg, K. (2008). The relevance of recent developments in classical conditioning to understanding the etiology and maintenance of anxiety disorders. Acta Psychologica, 127(3), 567–580. https://doi.org/10.1016/j.actpsy.2007.11.007

Ney, L. J., Wade, M., Reynolds, A., Zúj, D. V., Dymond, S., Matthews, A., & Felmingham, K. L. (2018). Critical evaluation of current data analysis strategies for psychophysiological measures of fear conditioning and extinction in humans. International Journal of Psychophysiology, 134, 95–107. https://doi.org/10.1016/j.ijpsycho.2018.10.010

Norrholm, S. D., Jovanovic, T., Olin, I. W., Sands, L. A., Karapanou, I., Bradley, B., & Ressler, K. J. (2011). Fear extinction in traumatized civilians with posttraumatic stress disorder: relation to symptom severity. Biological Psychiatry, 69(6), 556–563. https://doi.org/10.1016/j.biopsych.2010.09.013

Norrholm, S. D., Jovanovic, T., Smith, A. K., Binder, E., Klengel, T., Conneely, K., Mercer, K. B., Davis, J. S., Kerley, K., Winkler, J., Gillespie, C. F., Bradley, B., & Ressler, K. J. (2013). Differential genetic and epigenetic regulation of catechol-O-methyltransferase is associated with impaired fear inhibition in posttraumatic stress disorder. Frontiers in Behavioral Neuroscience, 7, 30. https://doi.org/10.3389/fnbeh.2013.00030

Orr, S. P., Metzger, L. J., Lasko, N. B., Macklin, M. L., Peri, T., & Pitman, R. K. (2000). De novo conditioning in trauma-exposed individuals with and without posttraumatic stress disorder. Journal of Abnormal Psychology, 109(2), 290–298. https://doi.org/10.1037/0021-843X.109.2.290

Pittig, A., Trenor, M., LeBeau, R. T., & Craske, M. G. (2018). The role of associative fear and avoidance learning in anxiety disorders: Gaps and directions for future research. Neuroscience and Biobehavioral Reviews, 88, 117–140. https://doi.org/10.1016/j.neubiorev.2018.03.015

Purves, K. L., Constantinou, E., McGregor, T., Lester, K. J., Barry, T. J., Trenor, M., Sun, M., Margraf, J., Craske, M. G., Breen, G., & Eley, T. C. (2019). Validating the use of a smartphone app for remote administration of a fear conditioning paradigm. Behaviour Research and Therapy, 123, 103475. https://doi.org/10.1016/j.brat.2019.103475

Rabinak, C. A., Mori, S., Lyons, M., Milad, M. R., & Phan, K. L. (2017). Acquisition of CS-US contingencies during Pavlovian fear conditioning and extinction in social anxiety disorder and posttraumatic stress disorder. Journal of Affective Disorders, 207, 76–85. https://doi.org/10.1016/j.jad.2016.09.018

Raeder, F., Merz, C. J., Magruf, J., & Zlomuzica, A. (2020). The association between fear extinction, the ability to accomplish exposure and exposure therapy outcome in specific phobia. Scientific Reports, 10(1), 4288. https://doi.org/10.1038/s41598-020-61004-3

Rimfeld, K., Malanchini, M., Sparo, T., Spickernell, G., Selzam, S., McMillan, A., Dale, P. S., Eley, T. C., & Plomin, R. (2019). Twins early development study: A genetically sensitive investigation into behavioral and cognitive development from infancy to emerging
adulthood. Twin Research and Human Genetics, 22(6), 508–513. https://doi.org/10.1017/thg.2019.56

Sijbrandij, M., Engelhard, I. M., Lommen, M. J. J., Leer, A., & Baas, J. M. P. (2013). Impaired fear inhibition learning predicts the persistence of symptoms of posttraumatic stress disorder (PTSD). Journal of Psychiatric Research, 47(12), 1991–1997. https://doi.org/10.1016/j.jpsychires.2013.09.008

Spitzer, R. L., Kroenke, K., Williams, J. B. W., & Löwe, B. (2006). A brief measure for assessing generalized anxiety disorder: The GAD-7. Archives of Internal Medicine, 166(10), 1092–1097. https://doi.org/10.1001/archinte.166.10.1092

Waters, A. M., & Pine, D. S. (2016). Evaluating differences in Pavlovian fear acquisition and extinction as predictors of outcome from cognitive behavioural therapy for anxious children. Journal of Child Psychology and Psychiatry, and Allied Disciplines, 57(7), 869–876. https://doi.org/10.1111/jcpp.12522

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