Is Body Mass Index the Prognostic Factor in Breast Cancer? : A Meta-Analysis

This study was performed to integrate the results of previous studies that investigated the relationship between body mass index (BMI) and prognosis in breast cancer. We reviewed the English literatures using the MEDLINE database from 1966 to 1999. The materials included 12 published articles with a total of 8,029 cases of breast cancer. The effect size was obtained from hazard ratio in each study. Homogeneity test was conducted before the integration of each effect size and the result demonstrated that the studies were heterogeneous. A random effect model was used to integrate the overall effect size. The integrated effect size was 1.56 (95% confidence interval, 1.22-2.00). In addition, publication bias should be accounted for because each published study was asymmetric in shape revealed by funnel plot. These results suggest that BMI have a prognostic significance in breast cancer. We believe that well-designed longitudinal studies, involving a large number of samples are required to resolve these issues.

Key Words : Breast Neoplasms; Prognosis; Body Mass Index; Meta-Analysis

INTRODUCTION

Breast cancer is known to be one of the most common female cancers in industrialized countries around the world. Also, the number of women suffering from breast cancer in developing countries has recently been increasing (1). Breast cancer is also the most common cause of death in cancer patients (2). In Korea, it ranks as the 3rd most common female cancer (3-5) and its mortality rate is 4.2 per 100,000 women, which accounted for more than 970 deaths in 1998 (6).

It has been reported that the factors that affect the prognosis of breast cancer after diagnosis include race (7, 8), age (7, 9-12), menopause (9, 13-15), period of disease (16), metastasis to the axillary lymph nodes (11, 17), estrogen receptors (18, 19), and weight, etc. Body size may also be related to prognosis in women with breast cancer. This relationship was first reported by Abe et al. (20), who studied obesity in Japanese women with breast cancer. After this several other investigators have studied the effect of body size on prognosis in breast cancer. Many investigators have found adverse effects associated with increased body size, whereas others have reported no effects (21). Commonly, in such studies body weight and body mass index (BMI) have been chosen as variables that represent body size.

The purpose of this study was to integrate the relationship between BMI, which is an indicator of body size, and prognosis in breast cancer by meta-analysis.

MATERIALS AND METHODS

Data collection

The materials used in this study were articles concerning the relationship between BMI and prognosis in breast cancer that were published in English between January 1966 and December 1999. Information was retrieved from the MEDLINE (only English publications) database. In addition, the cited references in published articles were also reviewed for relevant results. The medical subject headings for this search were breast cancer, prognosis, survival, body size, BMI,
and Quetelet’s index.

Meta-analysis

A total of 22 articles, which contained the above information, were selected and evaluated thoroughly by two reviewers. The agreement rate for the quality evaluation of each article between the two reviewers was 87.4%, and disagreements were adjusted by mutual consensus. The criteria for excluding articles for quantitative meta-analysis were as follows: i) articles that only compared survival rates versus BMI categories or ii) articles that researched the association between patient death without follow-up and BMI. In particular, the following standards were used for assessing the inclusion criteria of articles before applying the meta-analysis, and those that did not meet the standards were excluded from the analysis, even though they provided information for integrating effect size. Firstly, only original articles were used for the evaluation. Secondly, the investigators should have confirmed patient deaths with follow-up. Thirdly, the hazard ratio (effect size) and the 95% confidence interval (CI) must be included as the information about the relationship between BMI and prognosis in breast cancer, and be determined using Cox’s proportional hazard model. Adopting these criteria allowed twelve studies to be included in the meta-analysis.

The following information was collected from each study: the name of the first author, the year of publication, the country, the number of subjects, menopause state, definition of BMI including cut-off, effect size, 95% CI, etc. Particularly, the category of BMI in 12 studies was usually partitioned into binary, tertile, quartile, or quintiles. Therefore, we selected articles that were divisible into binary categories, to compare only the highest and lowest level of BMI. This method addresses only the question of the risk ratio between the extreme categories of exposure.

If the hazard ratio of breast cancer prognosis associated with the BMI was expressed in more than one way, the estimate extracted from the study was the one that reflected the greatest degree of controlling for confounders. Because few studies provided complete data for pre- and post-menopausal women separate, we chose the hazard ratio for the whole group if available.

The relationship between BMI and prognosis in breast cancer was assessed by the estimation of the effect size and 95% CI. In three studies (22-24), only the hazard ratio and the p-value were described, so we computed 95% CI of the hazard ratio by Greenland’s method (25).

To estimate the overall effect size, fixed-effects model and random-effects model were used. A general variance-based method using the confidence intervals based on the assumption of fixed effects was used for calculating the overall effect size, 95% CI, and Q statistics by Greenland (25), Prentice and Thomas (26). Because heterogeneity was assessed by the homogeneity test, we presented the overall effect size using the random-effects model as developed by DerSimonian and Laird (27). In addition, the sample sizes and effect size of each article were illustrated using funnel plot to assess publication bias (28).

RESULTS

Table 1 presents the details of the 12 studies used in the meta-analysis. A total of 8,029 breast cancer cases were included in the analysis. Six studies were performed in North America, five were published each in Europe and one in Asia. Two articles were written in the late 1980s and ten were written after the 1990s. In one paper (23), the subjects were women in pre-menopause. In another (29), the subjects were in post-menopause. In the remaining studies, the analyses of subjects were made regardless of the menopausal status. Each study was completed at the time of death. In 4 studies (8, 22, 34, 35), the only cause of deaths was breast cancer.

| First author       | Country  | No. of subjects | Menopausal status | BMI* Cut-off | Hazard ratio (death) |
|--------------------|----------|----------------|-------------------|--------------|----------------------|
| Greenberg (1985)   | UK       | 582            | pre               | >27.0 vs <20.0 | 1.8 (0.89-3.64)      |
| Mohle-Boetani (1988)| USA      | 838            |                   | ≥26.2 vs ≥22.9 | 1.4 (1.11-1.49)      |
| Kyogoku (1990)     | Japan    | 213            |                   | ≥25 vs <20    | 3.89 (0.77-19.70)    |
| Tornberg (1993)    | Sweden   | 1,170          |                   | ≥28 vs <22    | 1.7 (1.2-2.3)        |
| Eley (1994)        | USA      | 1,130          |                   | 85th vs 50th%| 2.5 (1.8-3.4)        |
| Holmberg (1994)    | Norway   | 359            |                   | ≥29 vs <19   | 5.93 (1.98-17.8)     |
| Jain (1994)        | Canada   | 1,033          | post              | >27.3 vs <22 | 0.78 (0.48-1.22)     |
| Tonkelaar (1995)   | Netherlands| 241         |                   | ≥26 vs <26   | 0.95 (0.51-1.78)     |
| Zhang (1995)       | USA      | 698            |                   | >8.8 vs <24.6| 1.9 (1.0-3.7)        |
| Maehle (1996)      | Norway   | 1,238          |                   | 5th vs 1st quintile| 1.37 (0.99-1.90)  |
| Galanis (1998)     | USA (Hawaii)| 378      |                   | ≥25.8 vs ≤22.6| 2.2 (0.9-5.4)        |
| Saxe (1999)        | USA      | 149            |                   | >27 vs ≤27   | 0.74 (0.32-1.71)     |

*BMI: body mass index.
For this study, we selected death as a standard for analysis regardless of the cause of death.

A homogeneity test was performed to check the results of each study. The statistical Q-value was 30.88, which is larger than the critical value of 19.67 when the degree of freedom is 11. No homogeneity was found, and therefore, all 12 studies were heterogeneous.

Fig. 1 shows the effect sizes and 95% CI of each study, and the overall effect size determined by meta-analysis. Three of the twelve studies only reported that the effect size was less than 1 and the remaining nine studies reported that the effect size was greater than 1. However, only five studies mentioned the statistical significance of these results. Using the random-effects model (27), which accounted for heterogeneity across the 12 studies, estimates based on an overall effect size for BMI and prognosis in breast cancer found to be 1.56 (95% CI, 1.22-2.00), which was statistically significant (Table 2).

Inverse funnel plotting between sample size and the effect sizes for each study is shown in the log scale (Fig. 2). The scatter of points showed that the studies were concentrated at the right lower and middle part of the triangle. Therefore, studies with a large number of breast cancer patients were lacking.

**DISCUSSION**

Studies on obesity and prognosis in breast cancer have been undertaken since the late 1970s. In this study, a meta-analysis was performed to integrate the relationship between BMI as an index of obesity and prognosis in breast cancer. We observed that high BMI had a bad affect on the prognosis in breast cancer with a statistical significance in this study.

However, because the information used in this study was based on observational studies, the confounding effect of each of the study variables should be accounted for when estimating the relationship between BMI and the prognosis of a breast cancer patient. Also, because the terms of reference of BMI used in each study were different, we were presented with limitations in the terms of explaining this result. Only the highest and lowest hazard ratio of BMI were integrated, and therefore, this study does not purport the relationship between BMI and prognosis in breast cancer across the BMI range. For these reasons, we thought that the strength of effect size did not mean the absolute value, while it did suggest the possibility of BMI as a prognostic factor related to the prognosis of breast cancer.

Subjects of each study were the patients who had been listed as having breast cancer in their community cancer registry, who had had a diagnosis at a defined group and were followed up prospectively, who had had medical treatment at specific hospitals, or who had been assigned to patient groups in a case-control study. Different methods of subject selection might also have influenced the results of this study. Since six studies were published in North America, five in

### Table 2. The overall effect size and 95% confidence interval for body mass index and prognosis in breast cancer according to statistical model

| Model        | Overall effect size | 95% confidence interval |
|--------------|---------------------|-------------------------|
| Fixed effect | 1.54                | 1.35-1.75               |
| Random effect| 1.56                | 1.22-2.00               |

**Fig. 1.** Effect sizes, overall effect size, and their 95% confidence intervals for body mass index and prognosis in breast cancer by study.

**Fig. 2.** Funnel plotting of sample size and the logarithm of the hazard ratio of body mass index and prognosis in breast cancer by study.
Europe, and one in Asia, it is possible that racial and environmental factors might also have influenced the results. In addition, an inconsistency might have existed in terms of the definition of death and the time of the final observation. In this study, there was a lack of clarity in terms of the recognition of the termination of study, and we will emphasize the importance of this particular aspect in further study.

We integrated the results of the studies chosen, which were based on effect size and 95% CI for each study using the fixed effects model and the random effects model. When the fixed effects model is used to integrate the results of studies, a homogeneity test is necessary (27). The hypothesis that the magnitude of the effect size was homogeneous was not accepted, and it was shown that the results of the studies incorporated into this analysis were heterogeneous. Because of diversity in the design and analysis of the various studies, we assumed that the effect size being estimated would vary among the studies. For example, due to different study population (nations, menopausal status, etc.) and the BMI cut-offs for the comparisons the true effects would vary. To account for several sources of variation in this study, DerSimonian and Laird had employed a random effects model to take account of the variation in the true effects of the studies being combined. Rather than rely on tests of homogeneity to establish the validity of the analysis, we assumed heterogeneity and employed the random effects models (Table 2) as developed by DerSimonian and Laird (27).

The findings of some meta-analyses have later been contradicted by large randomized controlled trials (38). Such discrepancies have brought discredit to a technique that has been controversial since its outset. The avoidance of misinterpretation of the results and publication bias was assessed using funnel plots. Funnel plot asymmetry is shown in areas with many subjects and large effect size (Fig. 2). We do not believe publication bias to be the source of asymmetry in this study. Rather, we think that the asymmetry may have been caused by the lack of the incorporation of larger-scale studies. As an evidence, although studies on the relationship between obesity and the prognosis in breast cancer have been conducted since the late 1970s (21), articles have only been published from 1990.

The primary targets of meta-analysis are randomized clinical trials in which there is little possibility of a bias intervening in the individual research project. Thus, a number of drawbacks are indicated in the application of this method to observational researches, and a considerable debate continues as to the merits of meta-analysis in observational studies (43-45). Nonetheless, many researchers have emphasized the significance of meta-analysis for contributing to knowledge by studying features across studies that may account for result variations (46-48).

It is still difficult to find a large cohort study of high quality targeted at women in this area. Therefore, what we should emphasize is not just a simple methodological application of meta-analysis, but the necessity to conduct high quality cohort studies and higher quality studies related to this issue. Only when high quality studies are performed, will quantitative meta-analysis contribute to the integration of high quality results.

In conclusion, our results suggest that BMI may be a prognostic indicator in breast cancer.

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