Stochastic Ordering of Infinite Geometric Galton-Watson Trees.

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Abstract

We consider Galton-Watson trees with Geom(p) offspring distribution. We let $T_\infty(p)$ denote such a tree conditioned on being infinite. We prove that for any $1/2 \leq p_1 < p_2 \leq 1$, there exists a coupling between $T_\infty(p_1)$ and $T_\infty(p_2)$ such that $\mathbb{P}(T_\infty(p_1) \subseteq T_\infty(p_2)) = 1$.

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1 Introduction

Consider a Galton-Watson tree $T(p)$ with some offspring distribution $\mu_p$ depending on the parameter $p$. If $p_1 < p_2$, then for many offspring distributions, it is possible to couple the trees $T(p_1)$ and $T(p_2)$ such that $\mathbb{P}(T(p_1) \subseteq T(p_2)) = 1$. This is for example the case with binomial, Poisson and negative binomial offspring distributions. One can then ask whether this property is preserved under certain conditioning. For example, if we let $T_\infty(p)$ denote a sample of $T(p)$ conditioned on being infinite, is it the case that there exists a coupling of $T_\infty(p_1)$ and $T_\infty(p_2)$, where $p_1 < p_2$, such that

$$\mathbb{P}(T_\infty(p_1) \subseteq T_\infty(p_2)) = 1? \quad (1)$$

Of course, if $p$ is such that the tree $T(p)$ is finite almost surely, then one needs to take care in defining $T_\infty(p)$.

It is known (see Example 1.1 in [1]) that there exists a parameterized offspring distribution $\mu_p$ such that $T(p_1) \subseteq T(p_2)$ can be made to hold almost surely, but for which (1) fails. However, [1] is known to hold for the Poisson offspring distribution (see Lyons, Peled and Schramm in [4]) and for the Binomial(d, p) offspring distribution (see [1]) when $d = 2, 3$. It would of course be desirable to give a complete characterization of the properties of $\mu_p$ needed to have (1). Unfortunately, this seems to be out of reach at the moment and the goal of this paper is to give another piece of the puzzle by studying the case of geometric offspring distribution (i.e. when for some $p \in (0,1)$, the probability that an individual has $k$ offspring is $p^k(1-p)$). Similar questions can be asked about percolation clusters on graphs. One of the most interesting cases is the corresponding problem for bond-percolation on $\mathbb{Z}^d$; see Open problem 4.3 of [1].

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Another similar type of question was studied by Luszczak and Winkler in [5]. The authors considered trees $T_k$, which are trees conditioned on being of a fixed size $k$, and asked the natural question whether these conditioned trees could be coupled in increasing order, i.e. so that $P(T_k \subset T_{k+1}) = 1$ (see Theorem 1 below for a precise statement of this result). We will use their result as a key ingredient in the proof of our main result.

We proceed to give some definitions needed for the statements of the main results. Let $o$ be the root, and consider the sets $V_0 = \{o\}$, $V_n := \{1,2,\ldots\}^n$ and $V = \{o\} \cup \bigcup_{n=1}^{\infty} V_n$. For any vertex $v = (v_1, \ldots, v_n) \in V_n$, with $n \geq 1$, we define $v^- = (v_1, \ldots, v_{n-1}) \in V_{n-1}$, with the convention that if $n = 1$, then $v^- = \{o\}$. We will think of $v^-$ as being the parent of $v$. Furthermore if $v = (v_1, \ldots, v_n) \in V_n$, is such that $v_n \geq 2$, we let $v' = (v_1, \ldots, v_n - 1) \in V_n$ and otherwise we let $v'$ be undefined. We think of $v'$ as being the ‘youngest older sibling’ of $v$. Of course, such a vertex only exists if $v$ is not the first child of $v^-$. For any two elements $u = (u_1, \ldots, u_k)$ and $v = (v_1, \ldots, v_l) \in V \setminus \{o\}$, we let $(u, v) = (u_1, \ldots, u_k, v_1, \ldots, v_l)$ denote the concatenation of $u$ and $v$. For $i \geq 1$, we will allow a slight abuse of notation, by writing $(u, i)$ instead of $(u, (i))$ for $(u_1, \ldots, u_k, i)$. Furthermore, for any $u \in V$ we let $(u, o) = (o, u) = u$ so that in particular, $(o, o) = o$.

Consider a set of vertices $V \subset V$. $V$ induces a set of edges if we "place" an edge between $v$ and $v^-$ whenever $v, v^- \in V$. A tree $T$ is a then a connected graph consisting of a set of vertices $V(T) \subset V$ and the induced set of edges $E(T)$. Since $V(T)$ determines the graph completely, we will leave $E(T)$ implicit. A subtree of a tree $T$ is defined to be a connected subgraph of $T$. For any $v \in V$ and a tree $T$, we let $T^v$ denote the subtree (of $T$) with vertex set $V(T^v) := \{u \in V(T) : u = (v, w) \text{ for some } w \in V\}$.

Note that if $v \not\in V(T)$, we get that $V(T^v) = \emptyset$. Informally, $T^v$ is simply the tree consisting of $v$ and the descendants of $v$ that belongs to $T$. We also define $H(T^v) := \{w \in V : (v, w) \in V(T^v)\}$, which is simply a shift of $T^v$, mapping $v$ to $o$. For $i \geq 1$, we sometimes abuse notation and write $T^i$ instead of $T^{(i)}$. We will let $|T|$ denote the number of vertices of a tree $T$ and we call this the size of $T$.

We will let $F$ denote the set of trees $T$, with the following two properties. Firstly, $o \in V(T)$, and secondly, if $v \in V(T)$ then $v' \in V(T)$ if $v'$ exists. For $1 \leq k < \infty$, let $T_k$ be uniformly chosen among the trees $T \in F$, such that $|T| = k$. Let the distribution of $T_k$ be denoted by $P_k$. We let $c_k$ denote the number of trees $T \in F$ such that $|T| = k$. Thus, $c_1 = 1, c_2 = 1, c_3 = 2, c_4 = 5, \ldots$. The trees of $F$ are sometimes called the rooted, ordered trees, and it is well known that $(c_k)_{k \geq 1}$ are the Catalan numbers (see for instance [3], Exercise 6.19).

The random trees with Geometric offspring distribution that we consider in this paper are formally defined as follows. For $p \in (0,1)$, let $(X_v)_{v \in V}$ be an i.i.d. collection of random variables, indexed by $V$, such that $P(X_v = k) = p^k(1-p)$ for $k = 0,1,\ldots$. Then, let $V_0(T(p)) = \{o\}$ and inductively let

$$V_{n+1}(T(p)) = \bigcup_{v \in V_n(T(p))} \bigcup_{1 \leq i \leq X_v} \{(v, i)\},$$

for every $n \geq 0$. We then define $T(p)$ to be the tree with vertex set

$$V(T(p)) = \bigcup_{n=0}^{\infty} V_n(T(p)).$$
Observe that \( V_n(T(p)) \subset V_n \) so that this is the set of vertices of \( T(p) \) at distance \( n \) from the root. Of course, if \( X_v = 0 \) for some \( v \), the second union of (2) is over an empty set and so no descendants of \( v \) belongs to \( T(p) \). For a similar reason, if for some \( n \), \( V_n(T(p)) = \emptyset \), it follows that \( V_{n+1}(T(p)) = \emptyset \). We also note that \( T(p) \in \mathcal{F} \).

It is easy to see from the construction, that for any tree \( T \in \mathcal{F} \) of size \( k \), \( \mathbb{P}(T(p) = T) = p^{k-1}(1-p)^k \), and therefore

\[
\mathbb{P}(|T(p)| = k) = c_k p^{k-1}(1-p)^k. 
\] (3)

It follows that the distribution of \( T(p) \), conditioned on the event that \( |T(p)| = k \), is \( \mathcal{T}_k \) (in particular, it is independent of \( p \)). Sometimes, it will be convenient to think of the empty set as the tree of size 0, and then we will use the notation \( T_0 = \emptyset \). For every \( 1 \leq k \leq \infty \), we define \( \eta_k(p) := \mathbb{P}(|T(p)| = k) \). Note that we allow for \( k = \infty \), and that \( \sum_{k=1}^{\infty} \eta_k(p) + \eta_{\infty}(p) = 1 \), since we always start with the root so that \( \mathbb{P}(|T(p)| = 0) = 0 \).

For two trees \( T, T' \) we say that \( T \subseteq T' \) if \( V(T) \subseteq V(T') \). Note that it follows by definition that \( E(T) \subseteq E(T') \). The following theorem is due to [5]. It is not explicitly stated, but as they point out (p. 427), it follows from their argument.

**Theorem 1.1 (Luczak, Winkler)** There exists a coupling of \( (T_k)_{k \geq 1} \) (where \( T_k \sim \mathcal{T}_k \) for every \( 1 \leq k < \infty \)) such that

\[
\mathbb{P}(T_1 \subset T_2 \subset \cdots) = 1. 
\] (4)

**Remark:** It is proved in [2] that equation (1) does not hold for general offspring distributions.

It is well known, that for \( p > 1/2 \), \( \mathbb{P}(|T(p)| = \infty) > 0 \). For such \( p > 1/2 \), let \( T_\infty(p) \) denote a random tree whose distribution equals that of \( T(p) \), conditioned on the event \( |T(p)| = \infty \). Furthermore, with \( (T_k)_{k \geq 1} \) as in Theorem 1.1 we define

\[
T_\infty(1/2) := \bigcup_{k=1}^{\infty} T_k. 
\] (5)

For any \( p \geq 1/2 \), we let \( \mathcal{T}_\infty(p) \) denote the distribution of \( T_\infty(p) \). Since \( (T_k)_{k \geq 1} \) does not depend on the parameter \( p \), it may seem strange to use the notation \( T_\infty(1/2) \) (i.e. \( p = 1/2 \)). However, in Section 2 we will describe a way to sample trees with distribution \( T_\infty(p) \), and then we will see that \( \bigcup_{k=1}^{\infty} T_k \) naturally corresponds to the critical case \( p = 1/2 \).

We can now state our main theorem.

**Theorem 1.2** For any \( 1/2 \leq p_1 < p_2 \leq 1 \), there exists a coupling of \( T_\infty(p_1) \) and \( T_\infty(p_2) \) (where \( T_\infty(p_1) \sim \mathcal{T}_\infty(p_1) \) and \( T_\infty(p_2) \sim \mathcal{T}_\infty(p_2) \)) such that

\[
\mathbb{P}(T_\infty(p_1) \subseteq T_\infty(p_2)) = 1. 
\]

**Remark:** We prove the theorem by giving an explicit construction of the coupling.

As mentioned before, the corresponding result for Galton-Watson trees with Poisson offspring distributions was proved in [3], while the corresponding result for Bin\((d,p)\) offspring distribution was proved in [1] for \( d = 2,3 \).

We have the following corollary.
Corollary 1.3 For any $k$ and $p \geq 1/2$, there exists a coupling of $T_k$ and $T_\infty(p)$, (where $T_k \sim T_k$ and $T_\infty(p) \sim T_\infty(p)$) such that
\[ P(T_k \subset T_\infty(p)) = 1. \]

Remark: This is an immediate consequence of Theorem 1.1 and Theorem 1.2.

We will end this section by briefly discussing the difficulties involved in proving Theorem 1.2. In order to construct a coupling satisfying the statement of Theorem 1.2, it is useful to be able to generate (or sample) trees with distribution $T_\infty(p)$. A natural way to do this is to use a sequential procedure using conditional probabilities as follows (where we will use the easily established result, see Lemma 2.1 that $\eta_\infty(p) = (2p - 1)/p$ for $p > 1/2$). Informally, we construct $\bar{T}(p) \sim T_\infty(p)$, by in the first step letting $T^1(p)$ be infinite with probability $P(|T^1(p)| = \infty) = \eta_\infty(p)/\eta_\infty(p) = p$, and of size $k < \infty$ with probability $P(|T^1(p)| = k|T(p)| = \infty) = p\eta_k(p)$. In the second step, if $|T^1(p)| = \infty$, then we let $T^2(p)$ be infinite with probability $P(|T^2(p)| = \infty|T^1(p)| = |T(p)| = \infty) = p\eta_\infty(p)$, while if $|T^1(p)| = k < \infty$, we let $T^2(p)$ be infinite with probability $P(|T^2(p)| = \infty|T^1(p)| = k, |T(p)| = \infty) = \cdots = p$. Continuing in this way until one finds a subtree which is infinite, and later a subtree which is of size 0 (which marks the end of the procedure), one can produce a tree $\bar{T}(p) \sim T_\infty(p)$.

This is one of the most natural ways of sampling a tree with distribution $T_\infty(p)$. However, it is not possible to use this procedure to construct a coupling proving Theorem 1.2. This can be seen by first observing that any coupling of $T_\infty(p_1)$ and $T_\infty(p_2)$ satisfying $P(T_\infty(p_1) \subseteq T_\infty(p_2)) = 1$ must certainly satisfy $|T^1_n(p_1)| \leq |T^1_n(p_2)|$ for every $n$. With positive probability, we could have that $|T^1(p_1)| < \infty$ while $|T^1(p_2)| = \infty$. Then, the conditional probability that $|T^2(p_1)| = \infty$ is $p_1$, while the conditional probability that $|T^2(p_2)| = \infty$ is $p_2\eta_\infty(p_2)$. Of course, for some choices of $p_1, p_2$ we can have that $p_1 > p_2\eta_\infty(p_2)$, and so we do not get that $|T^2(p_1)| \leq |T^2(p_2)|$ with probability 1.

In a second attempt, one might try to remedy the problem of our first attempt by first determining which subtrees $T^1(p), T^2(p), \ldots$ should be infinite, and which should be finite. Then, one could proceed by coupling these subtrees so that if $T^i(p_1)$ is infinite than so is $T^i(p_2)$. However, one will then find that with positive probability, both $T^1(p_1)$ and $T^1(p_2)$ are finite, and then one would have to let $|T^1(p_1)| = k$ with probability
\[ P(|T^1(p_1)| = k|T(p)| < \infty) = \frac{\eta_k(p_1)}{1 - \eta_\infty(p_1)} = c_k p_i^{k-1}(1 - p_i) \frac{p_i}{1 - p_i} = \eta_k(1 - p_i). \] (6)

Because of (6), we see that conditioned on $T^1(p_1)$ and $T^1(p_2)$ both being finite, we cannot have that $|T^1(p_1)| \leq |T^1(p_2)|$ with probability 1. In fact, a canonical coupling will result in $|T^1(p_2)| \leq |T^1(p_1)|$.

Therefore, the key to proving Theorem 1.2 is to find a way to generate a tree with distribution $T_\infty(p)$ which has the desired monotonicity properties and does not fall into any of the traps described above. In Section 2 we give this procedure along with some preliminary results. We then use this in Section 3 to prove Theorem 1.2.

2 Generating a tree with distribution $T_\infty(p)$

We start this section by proving the following easy lemma, already used in the introduction.
Lemma 2.1 If \( p \leq 1/2 \), then \( \eta_{\infty}(p) = 0 \) and if \( p > 1/2 \), then \( \eta_{\infty}(p) = \frac{2p-1}{p} \).

Proof. We have that

\[
1 - \eta_{\infty}(p) = \sum_{k=0}^{\infty} \mathbb{P}(X_o = k)(1 - \eta_{\infty}(p))^k
\]

\[
= \sum_{k=0}^{\infty} p^k(1-p)(1 - \eta_{\infty}(p))^k = \frac{1-p}{1-p(1-\eta_{\infty}(p))}.
\]

Solving for \( \eta_{\infty}(p) \), we have two solutions, \( \eta_{\infty}(p) = 0 \) and \( \eta_{\infty}(p) = \frac{2p-1}{p} \). It is well known that the tree is supercritical iff \( p > 1/2 \) from which the result follows. \( \square \)

We proceed by describing a procedure that will generate a tree \( \tilde{T}(p) \). In Lemma 2.2 we show that this procedure is well defined. Lemmas 2.3 and 2.4 will then prove that \( \tilde{T}(p) \sim \mathcal{T}_{\infty}(p) \) for every \( p \geq 1/2 \), while Lemma 2.5 will provide us with the crucial monotonicity properties used to prove our main results.

However, before we give any details, we will explain some heuristics of the construction and the coupling. Consider therefore a tree with distribution \( \mathcal{T}_{\infty}(1/2) \). It is well known (see for instance [3] sections 5 and 7) that such a tree will consist of one single infinite path to which there are smaller trees attached. In fact, it is possible to prove (and indeed we do this in Lemma 2.3) that such a tree can be generated in the following informally described way. Start with the root \( o \), and sample \( X_1, X_2 \) independently from a Geometric distribution with parameter \( 1/2 \). Then, attach independent trees with distribution \( \mathcal{T}(1/2) \) to the vertices \((1), \ldots, (X_1)\). Proceed by attaching an independent tree with distribution \( \mathcal{T}_{\infty}(1/2) \) to \((X_1 + 1)\) and end the construction by attaching independent trees with distribution \( \mathcal{T}(1/2) \) to the vertices \((X_1 + 2), \ldots, (X_1 + X_2 + 1)\). This implies that the probability that the infinite subtree can be found at position \( l \) (i.e. belongs to the \( l \)th child of the root) equals \( \mathbb{P}(X_1 = l - 1) = 2^{-l} \).

In order to construct the coupling of \( \tilde{T}(1/2) \sim \mathcal{T}(1/2) \) and \( \tilde{T}(p) \sim \mathcal{T}(p) \) for \( 1/2 < p \) in the desired way, we start by finding the infinite subtree of \( \tilde{T}(1/2) \) as above. For this, we simply use a random variable \( X \) such that \( \mathbb{P}(X = l) = 2^{-l} \). Given this number \( l \), the above paragraph tells us that for \( \tilde{T}(1/2) \), the subtrees \( \tilde{T}^1(1/2), \ldots, \tilde{T}^{l-1}(1/2) \) should be i.i.d. \( \mathcal{T}(1/2) \). We then require the corresponding subtree \( \tilde{T}^l(p) \) to be infinite also for \( \tilde{T}(p) \). The conditional distributions of \( \tilde{T}^1(p), \ldots, \tilde{T}^{l-1}(p) \) are slightly more complicated, but the key is that they will be independent, which facilitates a coupling of \( \tilde{T}^1(1/2) \) and \( \tilde{T}^l(p) \) for \( 1 \leq l \leq l-1 \) such that \( \tilde{T}^1(1/2) \subset \tilde{T}^l(p) \). Of course, it is not apriori clear that it is possible to achieve the coupling \( \tilde{T}^1(1/2) \subset \tilde{T}^l(p) \) but this is addressed in Lemma 2.4. We note that if our construction would instead condition on the smallest number \( k \) such that \( \tilde{T}^k(p) \) is infinite, then the conditional distributions of \( \tilde{T}^1(p), \ldots, \tilde{T}^{k-1}(p) \) would be \( \mathcal{T}(1-p) \) as described in [6]. Thus, just as described in the Introduction, our coupling would fail.

Of course, we also have to address the subtrees \( \tilde{T}^{l+1}(1/2), \ldots, \tilde{T}^{l+1}(p), \ldots \) which is done in a similar, albeit slightly more complicated (as the conditional distributions of \( \tilde{T}^{l+1}(p), \ldots \) are dependent of each other and \( \tilde{T}^1(p), \ldots, \tilde{T}^l(p) \) way) to ensure that the end result \( \tilde{T}(1/2) \) \( \tilde{T}(p) \) indeed has distribution \( \mathcal{T}_{\infty}(1/2) \) \( \mathcal{T}_{\infty}(p) \). Finally, it will turn out that the construction will not only work to couple \( \tilde{T}(1/2) \) and \( \tilde{T}(p) \), but will in fact work for any \( 1/2 \leq p_1 < p_2 \leq 1 \).
Below, \(X\) will (as discussed) determine the position of the single infinite subtree for \(\tilde{T}(1/2)\). \(L_m(p)\) will determine the size of the subtree \(\tilde{T}^m(p)\) while \(m_0 - 1\) will be the total number of subtrees.

We now turn to the formal description. We let \(X\) be such that \(P(X = l) = 2^{-l}\) for every \(l \geq 1\), and \((U_i)_{i \geq 1}\) be an i.i.d. sequence of \(U[0,1]\) random variables which is also independent of \(X\). For \(p \geq 1/2\), we then define the random variables \(L_1(p), L_2(p), \ldots\) through the following procedure

1. For every \(m < X\) and \(1 \leq k < \infty\), we let \(L_m(p) = k\) if
   \[
   \sum_{l=1}^{k-1} 2p\eta_l(p) \leq U_m < \sum_{l=1}^{k} 2p\eta_l(p),
   \]
   while if \(U_m > \sum_{l=1}^{\infty} 2p\eta_l(p)\) we let \(L_m(p) = \infty\).

2. We let \(L_X(p) = \infty\).

3. In order to define \(L_m(p)\) for every \(m > X\), we proceed sequentially. Assume therefore that \(L_{m-1}(p)\) has been determined. We then define \(n_{\infty}(p,m) := |\{i \leq m-1 : L_i(p) = \infty\}|\) and let \(L_m(p) = 0\) if
   \[
   U_m < \frac{2^{n_{\infty}(p,m)} p(1-p)}{(2^{n_{\infty}(p,m)} - 2)p + 1},
   \]
   and for \(1 \leq k < \infty\), we let \(L_m(p) = k\) if
   \[
   \frac{2^{n_{\infty}(p,m)} p(1-p)}{(2^{n_{\infty}(p,m)} - 2)p + 1} + \sum_{l=1}^{k-1} p\eta_l(p) \leq U_m < \frac{2^{n_{\infty}(p,m)} p(1-p)}{(2^{n_{\infty}(p,m)} - 2)p + 1} + \sum_{l=1}^{k} p\eta_l(p),
   \]
   and otherwise we let \(L_m(p) = \infty\).

For our construction of \(\tilde{T}(p)\), we will only use the \(L_m(p)\) such that \(m < m_0 := \min\{m : L_m(p) = 0\}\). We have the following lemma.

**Lemma 2.2** By the above construction, for any \(m < l\), we have that \(P(L_m(p) = \infty | X = l) = p\eta_{\infty}(p)\). Furthermore, we have that for \(m > l\),

\[
P(L_m(p) = \infty | X = l, n_{\infty}(p,m) = n) = p\eta_{\infty}(p) \frac{(2^n - 1)p}{(2^n - 2)p + 1}.
\]

**Proof.** Using that \(\eta_{\infty}(p) = (2p - 1)/p\), we have that

\[
p\eta_{\infty}(p) + \sum_{k=1}^{\infty} 2p\eta_k(p) = 2p - 1 + 2p(1 - \eta_{\infty}(p)) = 2p - 1 + 2(1 - p) = 1,
\]

and
and furthermore, for any \( n \geq 1 \),

\[
\frac{2^n p(1-p)}{(2^n - 2)p + 1} + \frac{(2^n - 1)p}{(2^n - 2)p + 1} p\eta_\infty(p) + \sum_{k=1}^\infty p\eta_k(p)
\]

\[= \frac{2^n p(1-p)}{(2^n - 2)p + 1} + \frac{(2^n - 1)p}{(2^n - 2)p + 1} (2p - 1 + p(1 - \eta_\infty(p)))
\]

\[= \frac{p^2(2^{n+1} - 2 - 2^n) + p(-2^n + 1 + 2^n)}{(2^n - 2)p + 1} + 1 - p
\]

\[= p(2^n - 2) + 1 (2^n - 2) p + 1 + 1 - p = 1,
\]

proving the lemma. \( \Box \)

**Remark:** Note that this also proves that the procedure determining \((L_m(p))_{m \geq 1}\) is well defined, since all relevant (conditional) probabilities are positive and sum to one.

We now turn to the construction of \( \tilde{T}(p) \).

1. Let \( X, (L_m(p))_{m \geq 1} \) be defined as above.

2. Let \( (T_{\infty,m}(p))_{m \geq 1} \) be an i.i.d. sequence which is independent of the other random variables in this list, and where \( T_{\infty,m}(p) \sim T_\infty(p) \) for every \( m \geq 1 \).

3. Let for every \( m \geq 1 \), \((T_{k,m})_{k \geq 1}\), be a sequence of random variables such that \( T_{k,m} \sim T_k \) for every \( k \geq 1 \). Furthermore, let \((T_{k,m})_{k \geq 1}\) be independent for different \( m \) and independent of the other random variables in this list.

Recall that \( m_0 = \min\{ m : L_m(p) = 0 \} \), and define the tree \( \tilde{T}(p) \) by

\[
V(\tilde{T}(p)) = \{o\} \bigcup_{m=1}^{m_0-1} \bigcup_{v \in T_{L_m(p),m}} \{(m, v)\}.
\]

Here, we abuse the notation somewhat in that we write \( T_{L_m(p),m} \) instead of \( T_{L_m(p),m}(p) \) when \( L_m(p) = \infty \). Informally, the tree \( \tilde{T}(p) \) is constructed by starting with a root and then attaching trees of size \( L_m(p) \) at the vertex \( \{(m)\} \) for every \( m < m_0 \). One key property of the tree \( \tilde{T}(p) \) is that \( \tilde{T}(p) \sim T_\infty(p) \), which we prove next. Since the proofs in the cases \( p = 1/2 \) and \( p > 1/2 \) are completely different, we split the result into two lemmas. We point out that Lemma 2.3 can be shortened by using the technique of size-biased Galton-Watson trees at criticality (see [3], Chapter 7, and the references within). However, in order to keep the paper self-contained, we give a proof from first principles.

**Lemma 2.3** We have that \( \tilde{T}(1/2) \sim T_\infty(1/2) \).
Proof. Recall from (3) the definition of $T_{\infty}(1/2)$. We have from this that for any $1 \leq l \leq m$, $1 \leq k_1, \ldots, k_{l-1}, k_{l+1}, k_m < \infty$, and with $k_l = k_l(n) = n - (k_1 + \cdots + k_{l-1} + k_{l+1} + \cdots + k_m) - 1$,

$$
\mathbb{P}(|T_{\infty}^1(1/2)| = k_1, \ldots, |T_{\infty}^{l-1}(1/2)| = k_{l-1}, |T_{\infty}^l(1/2)| = \infty,
|T_{\infty}^{l+1}(1/2)| = k_{l+1}, \ldots, |T_{\infty}^m(1/2)| = k_m, |T_{\infty}^{m+1}(1/2)| = 0)
= \lim_{n \to \infty} \mathbb{P}(|T_n^1| = k_1, \ldots, |T_n^m| = k_m, |T_n^{m+1}| = 0)
= \lim_{n \to \infty} \mathbb{P}(|T_n^t(p)| = k_1, \ldots, |T_n^m(p)| = k_m, |T_n^{m+1}(p)| = 0||T(p)| = n)
= \lim_{n \to \infty} \prod_{i=1}^m \frac{p n_{i-1} (1-p)}{n_{i}} (1-p) = \lim_{n \to \infty} \prod_{i=1}^m \frac{c_i}{n_i} \prod_{1 \leq i \leq m : i \neq l} c_i(1-p)
= \lim_{n \to \infty} \prod_{i=1}^m \frac{c_i}{n_i} \prod_{1 \leq i \leq m : i \neq l} c_i(1-p)
$$

where we use (3) and $k := k_1 + \cdots + k_{l-1} + k_{l+1} + \cdots + k_m$. Observe that the choice of $p \in (0,1)$ is irrelevant (as discussed in the introduction), and that if $|T(p)| = n$, then \(\sum_{j=1}^{\infty} |T_j(p)| = n - 1\) which explains the definition of $k_l$. Furthermore, as in Lemma 2.1 of [1], we have that $\lim_{n \to \infty} c_{n-1}/c_n = 1/4$, so we conclude that

$$
\mathbb{P}(|T_{\infty}^1(1/2)| = k_1, \ldots, |T_{\infty}^{l-1}(1/2)| = k_{l-1}, |T_{\infty}^l(1/2)| = \infty,
|T_{\infty}^{l+1}(1/2)| = k_{l+1}, \ldots, |T_{\infty}^m(1/2)| = k_m, |T_{\infty}^{m+1}(1/2)| = 0)
= \frac{1}{4^{l+1}} \left( \prod_{1 \leq i \leq m : i \neq l} c_i \right)
= \frac{1}{4} \prod_{1 \leq i \leq m : i \neq l} c_i
= \frac{1}{4} \prod_{1 \leq i \leq m : i \neq l} \frac{1}{2} \eta_i(1/2).
$$

We continue by noting that

$$
\mathbb{P}(|\hat{T}^1(1/2)| = k_1, \ldots, |\hat{T}^{l-1}(1/2)| = k_{l-1}, |\hat{T}^l(1/2)| = \infty,
|\hat{T}^{l+1}(1/2)| = k_{l+1}, \ldots, |\hat{T}^m(1/2)| = k_m, |\hat{T}^{m+1}(1/2)| = 0)
= \mathbb{P}(L_1(1/2) = k_1, \ldots, L_{l-1}(1/2) = k_{l-1}, L_l(1/2) = \infty,
L_{l+1}(1/2) = k_{l+1}, \ldots, L_m(1/2) = k_m, L_{m+1}(1/2) = 0||X = l)\mathbb{P}(X = l)
= \left( \prod_{i=1}^{l-1} \eta_i(1/2) \right) \left( \prod_{i=l+1}^m \frac{1}{2} \eta_i(1/2) \right) \frac{1}{2^{l-1}} = \frac{1}{4} \prod_{1 \leq i \leq m : i \neq l} \frac{1}{2} \eta_i(1/2).
$$

This establishes that the joint distribution of the sizes of $T_{\infty}^1(1/2), T_{\infty}^2(1/2), \ldots$ is the same as that of $\hat{T}^1(1/2), \hat{T}^2(1/2), \ldots$. Furthermore, it is easy to see that conditioned on $|T_{\infty}(1/2)| = k_i$, the distribution of $H(T_{\infty}^i(1/2))$ is $T_k$. By construction of $\hat{T}(1/2)$, the same holds for $\hat{T}(1/2)$. Furthermore, by (5), $T_{\infty}(1/2)$ is the limiting distribution of $T_n \sim T_n$ when letting $n \to \infty$. Therefore, conditioned on $|T_{\infty}(1/2)| = \infty$, we have that $H(T_{\infty}^i(1/2)) \sim T_{\infty}(1/2)$. By construction of $\hat{T}(1/2)$, it holds that also $H(\hat{T}(1/2)) \sim T_{\infty}(1/2)$. We conclude that $T_{\infty}(1/2)$ and $\hat{T}(1/2)$ must have the same distribution. $\square$
Lemma 2.4 For every $p > 1/2$, we have that $\tilde{T}(p) \sim T_\infty(p)$.

Proof. Let $I, J \subset \{1, 2, \ldots\}$ be such that $I \neq \emptyset$, $I \cap J = \emptyset$, and $I \cup J = \{1, 2, \ldots, I + J\}$ where $I = |I|$ and $J = |J|$. Informally, $I = \{i_1, \ldots, i_J\}$ will be the set of children with an infinite number of descendants, while $J = \{j_1, \ldots, j_J\}$ will be the set of children with a finite number of descendants. Here, we have ordered the elements of $I, J$ so that $i_1 < i_2 < \cdots < i_I$ and $j_1 < j_2 < \cdots < j_J$. Using that $T_\infty(p)$ is a random tree with the same distribution as $T(p)$ conditioned on being infinite, we observe that for any $k = (k_1, \ldots, k_J) \in \{1, 2, \ldots\}^J$

$$
\Pr(|T_{\infty}^k(p)| = \cdots = |T_{\infty}^k(p)| = \infty, |T_{\infty}^k(p)| = k_1, \ldots, |T_{\infty}^k(p)| = k_J, |T_{\infty}^k(p)| = 0) = \frac{(p\eta_\infty(p))^l}{\eta_\infty(p)} \prod_{j=1}^l p\eta_{k_j}(p)(1 - p) = p(1 - p)(p\eta_\infty(p))^{l-1} \prod_{j=1}^l p\eta_{k_j}(p). \tag{10}
$$

We now need to show that when using the construction of $\tilde{T}(p)$ we get the analogous expression. For any $i \in I$ let $j(i)$ be the smallest $j \in J$ such that $j > i$ if such a $j$ exists. We get that

$$
\Pr(|\tilde{T}_{\infty}^i(p)| = \cdots = |\tilde{T}_{\infty}^i(p)| = \infty, |\tilde{T}_{\infty}^i(p)| = k_1, \ldots, |\tilde{T}_{\infty}^i(p)| = k_J, |\tilde{T}_{\infty}^i(p)| = 0) =
$$

$$
= \sum_{I=1}^l 2^{-i} \Pr(L_{i_1}(p) = \cdots = L_{i_I}(p) = \infty, L_{j_1}(p) = k_1, \ldots, L_{j_J}(p) = k_J, L_{I+J+1}(p) = 0 | X = i)$$

$$
= \sum_{I=1}^l 2^{-i} (p\eta_\infty(p))^{l-1} \left( \prod_{j=m} p\eta_{k_m}(p) \right) \left( \prod_{j=m} p\eta_{k_m}(p) \right) 
$$

$$
\times \Pr(L_{i_{I+1}}(p) = \cdots = L_{i_{I+J+1}}(p) = \infty, L_{j_{i_1}}(p) = k_{j(i_1)}, \ldots, L_{j_J}(p) = k_J, L_{I+J+1}(p) = 0 | X = i_I). \tag{11}
$$

The first equality simply divides into cases depending on the value of $X$. The second equality uses (7) and the first part of Lemma 2.2. We proceed by considering (the possibly empty) set of $j$ such that $i_I < j < i_{I+1}$. Using that $2^{-i_I} \prod_{j_m < i_I} 2 = 2^{-I}$, we then get that (11) equals

$$
\sum_{I=1}^l 2^{-I} (p\eta_\infty(p))^{l-1} \left( \prod_{j=m} p\eta_{k_m}(p) \right) \left( \prod_{j=m} p\eta_{k_m}(p) \right) 
$$

$$
\times \Pr(L_{i_{I+1}}(p) = \cdots = L_{i_{I+J+1}}(p) = \infty, L_{j_{i_1}}(p) = k_{j(i_1)}, \ldots, L_{j_J}(p) = k_J, L_{I+J+1}(p) = 0 | X = i_I)
$$

$$
= \sum_{I=1}^l 2^{-l} (p\eta_\infty(p))^{l-1} \left( \prod_{j=m} p\eta_{k_m}(p) \right) p\eta_\infty(p) \left( \frac{2^l - 1}{2^l - 2} \right) \frac{p}{2^l - 2} 
$$

$$
\times \Pr(L_{i_{I+2}}(p) = \cdots = L_{i_I}(p) = \infty, L_{j_{i_1}}(p) = k_{j(i_1)}, \ldots, L_{j_J}(p) = k_J, L_{I+J+1}(p) = 0 | X = i_I),
$$

where we in the second step use Lemma 2.2. Iterating this procedure, and using (5), we get that

$$
\Pr(|\tilde{T}_{\infty}^i(p)| = \cdots = |\tilde{T}_{\infty}^i(p)| = \infty, |\tilde{T}_{\infty}^i(p)| = k_1, \ldots, |\tilde{T}_{\infty}^i(p)| = k_J, |\tilde{T}_{\infty}^i(p)| = 0) =
$$

$$
= \cdots = \left( \prod_{m=1}^l p\eta_{k_m}(p) \right) p\eta_\infty(p) \left( \frac{2^l - 1}{2^l - 2} \right) \frac{p}{2^l - 2}.$$
where
\[ \prod_{m=l}^{l-1} \frac{(2^m - 1)p}{(2^m - 2)p + 1} := 1, \quad (12) \]
if \( l = I \).

Comparing this to (10), we need to prove that for every \( I \),
\[ \sum_{l=1}^{I} 2^{-l} \left( \prod_{m=l}^{I-1} \frac{(2^m - 1)p}{(2^m - 2)p + 1} \right) \frac{2^l}{(2^l - 2)p + 1} = 1, \quad (13) \]
and this we do by induction. First, we note that this trivially holds for \( I = 1 \).
Assume therefore that it holds for \( I \geq 1 \), and observe that by (12),
\[ \sum_{l=1}^{I+1} 2^{-l} \left( \prod_{m=l}^{I-1} \frac{(2^m - 1)p}{(2^m - 2)p + 1} \right) \frac{2^{I+1}}{(2^{I+1} - 2)p + 1} \]
\[ = \frac{2^{I+1}}{(2^{I+1} - 2)p + 1} \left( 2^{-l+1} + \sum_{l=1}^{I} 2^{-l} \left( \prod_{m=l}^{I-1} \frac{(2^m - 1)p}{(2^m - 2)p + 1} \right) \frac{2^l}{(2^l - 2)p + 1} \right) \]
\[ = \frac{2^{I+1}}{(2^{I+1} - 2)p + 1} \left( 2^{-(l+1)} + \sum_{l=1}^{I} 2^{-l} \left( \prod_{m=l}^{I-1} \frac{(2^m - 1)p}{(2^m - 2)p + 1} \right) \frac{2^l}{(2^l - 2)p + 1} \right) = 1 + 2^{-(l+1)} \frac{2^l}{(2^{l+1} - 2)p + 1} = 1, \]
where we use the induction assumption in the third equality. \( \square \)

In order to prove Theorem 1.2 using our construction of \( \tilde{T}(p) \), we will need the monotonicity properties stated in our next lemma, which consists of three parts.

**Lemma 2.5** Consider the functions
\[ p_{\eta_k}(p), \quad (14) \]
and
\[ \frac{2^n p(1-p)}{(2^n - 2)p + 1}. \quad (15) \]

1. For any \( 1 \leq k < \infty \), the function of (14) is non-increasing in \( p \) for \( p \geq 1/2 \).
2. For any \( n \geq 1 \), the function of (15) is non-increasing in \( p \) for \( p \geq 1/2 \).
3. For any \( 1/2 \leq p < 1 \), the function of (15) is non-increasing in \( n \) for \( n \geq 1 \).

**Proof.** We have that \( p_{\eta_k}(p) = c_k p^k (1-p)^k \), which is clearly non-increasing in \( p \) for \( p \geq 1/2 \). Furthermore, since \( p(1-p) \) and \( 1/((2^n - 2)p + 1) \) are non-increasing in \( p \) for \( p \geq 1/2 \) and any \( n \geq 1 \), the first two parts of the statement follows.

The third part follows by observing that the function of (15) is non-increasing in \( n \) if
\[ (2^{n+1} - 2)p + 1 \geq 2^{n+1} ((2^n - 2)p + 1), \]
which simplifies to \( 1 - 2p \geq 2 - 4p \) and holds for all \( p \in [1/2, 1] \). \( \square \)
3 Proof of Theorem 3.1

In order to prove Theorem 3.1, we will start by proving it in a special case, and then use this to prove the full statement.

Theorem 3.1 For any $p > 1/2$, there exists a coupling of $T_\infty(1/2)$ and $T_\infty(p)$ (where $T_\infty(1/2) \sim T_\infty(1/2)$ and $T_\infty(p) \sim T_\infty(p)\)) such that

$$P(T_\infty(1/2) \subset T_\infty(p)) = 1.$$  

Remark: The proof of Theorem 1.2 will be very similar, and therefore we will only address the necessary changes.

Proof of Theorem 3.1. In order to facilitate the display of formulas, we will use the notation $p_1 = 1/2$ and $p_2 = p$.

We will construct a sequence of pairs of trees $(\tilde{T}_l(p_1), \tilde{T}_l(p_2))_{l \geq 1}$ such that $\tilde{T}_l(p_i) \sim T_\infty(p_i)$ for every $l \geq 1$, and

$$P(\tilde{T}_l(p_1) \cap (\bigcup_{n=0}^l \mathcal{V}_n) \subset \tilde{T}_l(p_2) \cap (\bigcup_{n=0}^l \mathcal{V}_n)) = 1,$$  

for every $l \geq 1$. That is, $(\tilde{T}_l(p_1), \tilde{T}_l(p_2))$ will be ordered up to distance $l$ from the root. From this, the statement will easily follow.

We start by proving (16) for $l = 1$. We will use the procedure of Section 2 to generate our trees, but we will do it simultaneously for $p_1$ and $p_2$. Consider therefore the following random variables that we will use in our construction of $(\tilde{T}_1(p_1), \tilde{T}_1(p_2))$.

1. $X$ is such that $P(X = l) = 2^{-l}$ for $l \geq 1$. It is also independent of all other random variables listed.

2. $(U_i)_{i \geq 1}$ is an i.i.d. collection of $U[0,1]$ random variables which is also independent of all other random variables listed.

3. For every $m \geq 1$, $(T_{k,m})_{k \geq 1}$ is a collection of random variables such that $T_{k,m} \sim T_k$ and $P(T_{1,m} \subset T_{2,m} \subset \cdots) = 1$, which is possible by Theorem 1.1. Furthermore, we take $(T_{k,m})_{k \geq 1}$ to be independent for different $m$, and independent from all other random variables listed. We also let $T_{\infty,m}(p_1) = \bigcup_{k=1}^\infty T_{k,m}$.

4. Finally, $(T_{\infty,m}(p_2))_{m \geq 1}$ is an i.i.d. collection such that $T_{\infty,m}(p_2) \sim T_\infty(p_2)$ for every $m$, and $(T_{\infty,m}(p_2))_{m \geq 1}$ is independent from all other random variables listed.

For any $m < X$ and $1 \leq k < \infty$, we let $L_m(p_i) = k$ iff

$$\sum_{i=1}^{k-1} 2p_i \eta(p_i) \leq U_m \leq \sum_{i=1}^k 2p_i \eta(p_i)$$

and otherwise we let $L_m(p_i) = \infty$. Thus, we see that $L_m(p_i)$ is chosen with probabilities as in (7) and in Lemma 2.2. Furthermore, by the first part of Lemma 2.5 we conclude that
\(L_m(p_1) \leq L_m(p_2)\) for every such \(m\). We also let \(L_X(p_1) = L_X(p_2) = \infty\). Recall that \(n_\infty(p_i, m) = |\{j \leq m - 1 : L_j(p_i) = \infty\}|\) and note that by our construction,

\[n_\infty(p_1, X) \leq n_\infty(p_2, X).\]  

(17)

We proceed as follows. For \(m = X + 1\), we let \(L_m(p_i) = 0\) iff

\[U_m < \frac{2^{n_\infty(p_i, m)} p_i(1 - p_i)}{(2^{n_\infty(p_i, m)} - 2)p_i + 1},\]

and for \(1 \leq k < \infty\), we let \(L_m(p_i) = k\) iff

\[\frac{2^{n_\infty(p_i, m)} p_i(1 - p_i)}{(2^{n_\infty(p_i, m)} - 2)p_i + 1} + \sum_{i=1}^{k-1} p_i \eta_i(p_i) \leq U_m < \frac{2^{n_\infty(p_i, m)} p_i(1 - p_i)}{(2^{n_\infty(p_i, m)} - 2)p_i + 1} + \sum_{i=1}^{k} p_i \eta_i(p_i),\]

and otherwise we let \(L_m(p_i) = \infty\). By Lemma 2.3, \(L_m(p_1) \leq L_m(p_2)\) and in particular it is possible for \(L_m(p_1) = 0\) while \(L_m(p_2) > 0\). This uses all three parts of that lemma. As a consequence, we get that \(n_\infty(p_1, X + 1) \leq n_\infty(p_2, X + 1)\).

Continue for \(m = X + 2\) etc in the natural way, and define \(m_0(p_i) = \min\{m : L_m(p_i) = 0\}\). As above, Lemma 2.3 shows that \(L_m(p_1) \leq L_m(p_2)\) for every \(m\). Define

\[V(\tilde{T}_1(p_i)) := \{0\} \bigcup_{m=1}^{m_0(p_i)-1} \bigcup_{v \in T_{L_m(p_i), m}} \{(m, v)\},\]  

(18)

where again, we abuse notation by writing \(T_{L_m(p_i), m}\) instead of \(T_{L_m(p_i), m}(p_i)\) when \(L_m(p_i) = \infty\).

Thus, \(\tilde{T}_1(p_i)\) is defined through the procedure of Section 2, and so by Lemmas 2.3 and 2.4 we see that \(\tilde{T}(p_i) \sim T\) for \(i = 1, 2\). By construction, we have that

\[P(\tilde{T}_1(p_1) \cap \{(0) \cup V_1\} \subset \tilde{T}_1(p_2) \cap \{(0) \cup V_1\}) = 1.\]

We now proceed inductively, and so assume that (16) holds for some \(l \geq 1\). We repeat the above procedure, but using 3' below instead of points 3 and 4.

3' For every \(m \geq 1\), \((T_{k,m})_{k \geq 1}\) is a collection of random variables such that \(T_{k,m} \sim T_k\) and

\[P(T_{1,m} \subset T_{2,m} \subset \cdots) = 1,\]

which is possible by Theorem 1.1. We take \((T_{k,m})_{k \geq 1}\) to be independent for different \(m\), and also independent from all other random variables listed, and let \(T_{\infty,m}(p_1) = \bigcup_{k=1}^{\infty} T_{k,m}\). Furthermore, we let \((T_{\infty,m}(p_2))_{m \geq 1}\) be such that \(T_{\infty,m}(p_2) \sim T_{\infty}(p_2)\) and coupled with \(T_{\infty,m}(p_1)\) so that

\[P(T_{\infty,m}(p_1) \cap (\cup_{n=0}^{l} V_n) \subset T_{\infty,m}(p_2) \cap (\cup_{n=0}^{l} V_n)) = 1.\]

This is possible by the induction hypothesis.

We then construct \((\tilde{T}_{l+1}(p_1), \tilde{T}_{l+1}(p_2))\) using 1,2 and 3'. To see why \((\tilde{T}_{l+1}(p_1), \tilde{T}_{l+1}(p_2))\) satisfies (16) for \(l + 1\), it suffices to observe the following. By the induction hypothesis, any pair
of trees used in [18] are already ordered up to distance \( l \) from their roots, and so by attaching
them to vertices at distance one from the root of \( \tilde{T}_{l+1}(p_i) \), the new trees are ordered up to
distance \( l + 1 \) from their roots.

In order to conclude the argument, let \( \gamma_l \) be a measure on \( \{0, 1\}^V \times \{0, 1\}^V \) with marginal
distributions \( T_\infty(p_1) \) and \( T_\infty(p_2) \) such that \( \gamma_l(\xi(\cup_{n=0}^l V_n) \leq \eta(\cup_{n=0}^l V_n)) = 1 \). Here, we identify
a tree \( T \) and an element \( \xi_T \in \{0, 1\}^V \) by letting \( \xi_T(v) = 1 \) iff \( v \in T \), and the measure \( \gamma_l \)
exists by the above construction. Since \( \{0, 1\}^V \times \{0, 1\}^V \) is compact, there exists a subsequen-
tial limiting measure \( \gamma \) on \( \{0, 1\}^V \times \{0, 1\}^V \) with marginal distributions \( T_\infty(p_1) \) and \( T_\infty(p_2) \) such that
\( \gamma(\xi(V) \leq \eta(V)) = \lim_{l \to \infty} \gamma_l(\xi(\cup_{n=0}^l V_n) \leq \eta(\cup_{n=0}^l V_n)) = 1 \). By Strassen’s theorem
it follows that there exists random trees \( T_\infty(p_1) \sim T_\infty(p_1) \) and \( T_\infty(p_2) \sim T_\infty(p_2) \) such that
\( \mathbb{P}(T_\infty(p_1) \subset T_\infty(p_2)) = 1 \).

We are now ready to prove Theorem 1.2. Since the proof is very similar to the proof of
Theorem 3.1, we will only address the minor changes.

**Proof of Theorem 1.2** We can assume that \( p_1 > 1/2 \) since the case \( p_1 = 1/2 \) has already
been proved. We change point 3 to

3 For every \( m \geq 1 \), \( (T_{k,m})_{k \geq 1} \) and \( T_{\infty,m}(p_1) \) is a collection of random variables such that
\( T_{k,m} \sim T_k, T_{\infty,m}(p_1) \sim T_\infty(p_1) \) and coupled so that \( \mathbb{P}(T_{1,m} \subset T_{2,m} \subset \cdots \subset T_{\infty}(p_1)) = 1 \),
which is possible by Theorems 1.1 and 3.1. Furthermore, we take \( (T_{k,m})_{k \geq 1} \) and \( T_{\infty,m}(p_1) \)
to be independent for different \( m \), and also independent from all other random variables
listed.

Point 3’ is changed accordingly.

\( \square \)

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