Forecasting of a Hierarchical Functional Time Series on Example of Macromodel for the Day and Night Air Pollution in Silesia Region — A Critical Overview

Daniel Kosiorowski∗, Dominik Mielczarek†, Jerzy P. Rydlewski‡

Submitted: 26.11.2017, Accepted: 23.02.2018

Abstract

In economics we often face a system which intrinsically imposes a structure of hierarchy of its components, i.e., in modeling trade accounts related to foreign exchange or in optimization of regional air protection policy. A problem of reconciliation of forecasts obtained on different levels of hierarchy has been addressed in the statistical and econometric literature many times and concerns bringing together forecasts obtained independently at different levels of hierarchy. This paper deals with this issue with regard to a hierarchical functional time series. We present and critically discuss a state of art and indicate opportunities of an application of these methods to a certain environment protection problem. We critically compare the best predictor known from the literature with our own original proposal. Within the paper we study a macromodel describing the day and night air pollution in Silesia region divided into five subregions.

Keywords: day and night air pollution, functional data analysis, functional median, hierarchical time series, reconciliation of forecasts

JEL Classification: C53, C32, R11

∗Cracow University of Economics
†AGH University of Science and Technology
‡AGH University of Science and Technology; e-mail: ry@agh.edu.pl
1 Introduction

A variety of economic systems consists of a certain class of subsystems, which form a fixed hierarchical structure, i.e.:

1. A country considered with respect to monthly Gross National Product with a splitting into regions and subregions
2. A national balance of trade with a splitting into branches and subbranches of industry, services and agriculture
3. Weekly total inflows and outflows of current accounts for a certain group of clients (priority, individual, corporate clients, or gender, or demographic groups) considered in 5-min consecutive intervals
4. A turnover of a company with regard to product lines and/or client target groups
5. Social and health care insurance costs divided into age and place of living segments
6. Social inequalities of households with regard to education level, religious faith, political choice or ethnicity
7. Social and health costs associated with environment pollution for a certain region divided into subregions

Both from a theoretical as well as from practical point of view it is especially important to find a reliable method of modeling and forecasting a time evolution of a system similar to the above systems exhibiting a hierarchical structure. The method should be computationally tractable.

The issue is very closely tied with the reconciliation of forecasts - a problem known from the econometric literature (see Shlifer and Wolff 1979, Kohn 1982, Weale 1988, Kahn 1998, and Fliedner 2001).

It is often observed that forecasts prepared for lower levels of hierarchy do not sum to forecasts prepared for upper levels and the top level of the system. That fact may be caused for example by an application of different measurement methodologies, different precision for different levels.

In this paper we focus our attention on the problem of forecasting a hierarchic system describing the day and night air pollution in Silesia region in Poland. The region is divided into five subregions.

The day and night air pollution is treated as a realization of a functional random variable. Hence we consider a forecasting of a hierarchical functional time series (HFTS). A predictor possessing good statistical properties in this case is directly connected with an opportunity of designing cost-effective pro-ecological regional politics, which optimize social welfare being a function of the day and night air pollution. It is worth stressing, that while using functional time series (FTS)
framework instead of well known one-dimensional time series setup, we forecast whole
day and night periods instead of predicting hour after hour. Note also, that using FTS
one can easily model and predict non-equally spaced time series, which may cause
fundamental problems for analysts using classical ARIMA, SARIMA methodology
(see Kosiorowski 2016, and Górecki et al. 2016). Our paper critically discusses the
best proposals known from the literature (see Shang and Hyndman 2017) and compare
them with our proposals.
The rest of the paper is organized as follows. Section 2 discusses the concept of
hierarchical time series. Section 3 discusses the concept of hierarchical functional
time series. Section 4 discusses the methods of HFTS forecasting. Section 5 contains
the empirical study – an example of a macromodel for the day and night air pollution
in Silesia region.

2 Hierarchical time series

Hierarchical time series (HTS) is a time series, where some fixed, often natural,
hierarchy is imposed. In other words, HTS can be considered as a time series, where
at each time we have insight to the values for any single variable at any level in the
structure with a fixed hierarchy. In Figure 1 an example visualization of hierarchical
time series at moment \( n \) is depicted. In the Figure 1, the observation made on the top
level is divided into two sublevels or level – 1 levels, and the observation made on the
level – 1 is divided into two level – 2 levels, but the division might be quite different,
and the only constraint is, that any level could be divided into finitely many number
of levels and the total number of observations is finite. Obviously, one can compute a
forecast for all series at all hierarchy levels independently, but the forecast at the lower
level do not sum to the forecast at the upper level. Hence, no reconciliation is made.
In the hierarchical setup the forecasting might be done in the following manners. The
forecast is made on the bottom level of the hierarchy. Subsequently, the aggregation
of the obtained forecast, basing on some historical data, is made on the upper level
of the hierarchy. This procedure is repeated upwards the hierarchy, until we get a
forecast on the top level. The method is called the bottom-up method. Conversely
we proceed in the top-down method, where the forecast is made on the top level.
The disaggregation is then made, so that we obtain a forecast on the lower levels
of the hierarchy. The methods are often mixed, as we obviously, for some reasons,
might be interested in the forecast on some intermediate level of the hierarchy. Then
the forecast is aggregated upward the hierarchy, and disaggregated downward the
hierarchy. The methods do not take into account the correlation structure of the
hierarchy. Prediction intervals for the forecasts are undefined as well. The more
detailed discussion and references the interested reader may find in Shlifer and Wolff
(1979), Kohn (1982), Weale (1988), Kahn (1998), and Fliedner (2001).
In their paper Hyndman et al. (2011) proposed a novel optimal combination forecast
method for HTS. Their proposal is based on independently forecasting all series at
all levels of the hierarchy and then using a regression model to obtain a reconciliated forecast. The forecasts obtained with their method add across the fixed hierarchy. It is also mean-unbiased and under some reasonable assumptions has minimum variance among linear combination of independent forecasts. They represent the fixed hierarchical structure in the matrix form. This approach allows for the correlations between the series at each level of the hierarchy. However, they mention that some computational problems may occur. They are connected to the inverse of relatively large, sparse matrices and solution of sparse linear least squares problem. Nevertheless, Hyndman et al. (2011) approach enables to obtain a reconciliated forecast for a considered phenomenon reconciliated with individual forecasts obtained on different levels of hierarchy.

Figure 1: An example visualization of hierarchical time series

3 Functional hierarchical time series

Functional hierarchical time series is a series which consists of functional data, i.e. we consider a hierarchical dataset of functions instead of real numbers or vectors in \( \mathbb{R}^m \). The functional data methods are described in monographies of Ferraty and Vieu (2006), Ramsay et al. (2009), and Horváth and Kokoszka (2012). Some statistical
tests have been recently developed for the functional framework as well (e.g. Kosiorowski et al. 2017a). Note that the methods developed for the uni- or multivariate HTS and described in Section 2 cannot be directly applied in HFTS setup. Many theoretical problems arise here, but note, that even the order of functions cannot be measured as easily as in the univariate case. The naïve forecast is not convincing as well, because it does not take into account a time dependency. The pointwise average can be easily calculated, but we usually do not know the true distribution on the $L^2[0,T]$ space, from which our data come from, so we cannot straightforwardly assume that the functional expected value exists, which makes the approach unconvincing as well. For the same reason the pointwise moving average seems out of the question. In their paper Shang and Hyndman (2017) proposed their method of HFTS forecasting. Their approach originates from their previous study (Hyndman et al. 2011), described in Section 2, which takes into account the whole hierarchical structure of the data. The reconciliated forecast for a fixed hierarchical structure with $L$ levels takes a following shape

$$\hat{X}_{n+1}(t) = F(\hat{x}_{\text{level} subscription}(t), \hat{x}_{\text{level} 11}(t), \ldots, \hat{x}_{\text{level} 1i 1}(t), \ldots, \hat{x}_{\text{level} Li L}(t)), \ldots)$$

where $\hat{x}_{\text{level} ki}(t)$ denotes an $i_k$ forecast obtained for the functional time series at level $k$ and $F$ denotes a certain generalized least squares estimator. The HFTS structure at day $n$ is described by a matrix equation

$$X_n = S_n b_n, \quad (1)$$

where vector $X_n = (x_{\text{level} top}, x_{\text{level} 11}, \ldots, x_{\text{level} Li 1}, \ldots, x_{\text{level} Li L}, \ldots)$, that is, it is containing all series at all levels of hierarchy, $b_n$ is a vector representing the series at the lowest level of the fixed hierarchy, and $S_n$ is a finite matrix that shows the connection between the vectors $X_n$ and $b_n$. The forecast is made then, that is:

$$\hat{X}_{n+1} = S_{n+1} \hat{\beta}_{n+1} + \epsilon_{n+1},$$

where $\hat{X}_{n+1}$ is a matrix of forecasts made for all series at all levels of the fixed hierarchy, $\hat{\beta}_{n+1} = E[b_{n+1} | X_1, \ldots, X_n]$ is an unknown multivariate expected value of a forecast distribution for the most disaggregated series and $\epsilon_{n+1}$ represents the errors of the reconciliation. Note that the level forecasts are obtained using non-robust method, which maps functional time series into one dimensional series of functional component scores (for details see Kosiorowski 2014). Components $\hat{\beta}_{n+1}$ are estimated in the study by Shang and Hyndman (2017) with the generalized least squares method, i.e.

$$\hat{\beta}_{n+1} = \left(S_{n+1}^T V^{-1} S_{n+1} \right)^{-1} S_{n+1}^T V^{-1} \hat{X}_{n+1}.$$
where a diagonal matrix $V$ estimates variances of series forecasts. The final forecast stems from the equation

$$\hat{X}_{n+1} = S_{n+1}\hat{\beta}_{n+1}.$$ 

Shang and Hyndman (2017) method has some important advantages and disadvantages. The Shang and Hyndman (2017) forecasts are aggregate consistent – they satisfy an aggregation constrains and are mean-unbiased. However, the method is very computationally demanding and sophisticated due to the sparse linear least squares problem and necessity of computing generalized inverses of large and sparse matrices (see 1). The method may be therefore robustified (for details see Kosiorowski et al. 2017b).

Note that Shang and Hyndman’s (2017) approach is equipped with an outright internal mechanism of forecasts reconciliation. Our proposal – which is described further – the reconciliation of forecasts is a byproduct of a fact that modified band depth (MBD, see López-Pintado and Romo 2007, 2009) is nontransitive. Shang and Hyndman (2017) reduce the problem of functional data forecasting to functional principal component regression: functions are represented in an empirical principal components base, then they use Hyndman’s functional regression basing on one-dimensional stationary time series modeling (see Kosiorowski 2014) and the authors assume that residual functions are approximately stationary (we do not make the restriction).

The last analyzed approach comes from the paper Kosiorowski et al. (2017c). The authors presented double functional median method and compared their method with Shang and Hyndman’s (2017) method as a reference approach.

Modified band depth (MBD, see López-Pintado and Romo 2007, 2009) of curve $x$ with respect to functional sample $X^N$ (a sample of $N$ functions, i.e., $X^N = \{x_i(t), i = 1, 2, \ldots, N \}$ and $t \in [0, T]$), estimates the curves’ frequency of being in the center. Note, that Zuo and Serfling (2000) formulated general conception of statistical depth function and Nieto-Reyes and Battey (2016) have proved that a depth for functional data is correctly defined. Nevertheless, we have a sample of $N$ functions $x_1, \ldots, x_N$.

Firstly we need to define sets of the following form

$$A(x; x_{i_1}, x_{i_2}) = \{ t \in [0, T] : \min_{r=i_1,i_2} x_r(t) \leq x(t) \leq \max_{r=i_1,i_2} x_r(t) \}.$$

Consequently, MBD can be defined, a functional depth, which takes into account a proportion of “time”, when $x$ is in the band made with two functions, i.e.

$$MBD(x|X^N) = \frac{2}{N(N-1)} \sum_{1 \leq i_1 < i_2 \leq N} \frac{\lambda(A(x; x_{i_1}, x_{i_2}))}{\lambda([0, T])},$$

where $\lambda$ is a Lebesgue measure.

Subsequently, the nested regions for the chosen functional depth can be constructed, that is, consider $MBD(x|X^N) \geq \alpha$. The median with respect to the considered functional depth is the most central observation. We define a sample median as

$$MED_{MBD}(X^N) = \arg \max_{i=1, \ldots, N} MBD(x_i|X^N).$$

D. Kosiorowski et al. 58
CEJEME 10: 53-73 (2018)
Forecasting of a Hierarchical Functional Time Series...

If more than one function achieves the depth maximum value, the median is defined as the average of the curves maximizing depth. Then we use a moving functional median:

\[ \hat{x}_{n+1} = MED_{MBD}(W_{n,k}), \]

where \( W_{n,k} \) is a moving window of a length \( k \) with an end in a moment \( n \), that is, \( W_{n,k} = \{ x_{n-k+1}, \ldots, x_n \} \). Double functional median method can be described in the following steps:

1. We calculate the moving functional median related to the MBD or another functional depth for each unit at the lowest level of hierarchy, i.e., \( MED_{MBD}(W_{n,k}) \). In empirical example analyzed in Section 5 for each town and at moment \( n \), we compute a functional median from a moving window of length 10 with respect to the functional depth \( MBD \):

\[ \hat{x}_{town}^{n+1} = MED_{MBD}\{ x_n^{town}, x_{n-1}^{town}, \ldots, x_n^{town} \}. \]

2. We calculate for the lowest but one level of hierarchy, a functional median from the medians calculated in the first step.

3. We repeat the second step until we calculate the functional median for the top level of the hierarchy.

In our empirical example, the second step is the last one, and finally, we obtain a forecast for \( n = 10, \ldots, 181 \):

\[ \hat{x}_{n+1} = MED_{MBD}\{ \hat{x}_{n+1}^{town1}, \ldots, \hat{x}_{n+1}^{town5} \}. \]

Hierarchical structure of the data is taken into account in the process of computing functional median of lower level functional medians, as translation of a single functional observation into neighbouring unit alters the outcome (for details see Kosiorowski et al. 2017c).

4 A critical overview of HFTS approaches

In a general case, an uncertainty evaluation of the HFTS forecast is an open issue. Due to insufficient theoretical background for conducting a precise statistical inference, in our approach we decided to expand ideas indicated by López-Pintado et al. (2010). Shang and Hyndman (2017) has obtained a representation of functions in the \( L^2 \) space with a Fourier basis. The Fourier basis is adjusted to the functional data they consider, because the data they analyzed were expected to be periodical. Afterwards they transformed functional time series into a family of one-dimensional principal component scores series. Then a maximum entropy bootstrap methodology proposed by Vinod and de Lacalle (2009) and implemented in \textit{meboot} R package, which is appropriate for time series setting, has been used. Although this simplification of the
problem seems to be attractive, it divests an analyst of the richness of behaviors a functional time series in comparison to one dimensional time series. We recommend using functional boxplots and adjusted functional boxplots (one can focus on sizes of boxes and $\alpha$--central regions), which realizes an idea of bootstrap for functional time series for the rough evaluation of the forecast uncertainty (for details see López-Pintado et al. 2010 and Sun and Genton 2011). It does not make much sense to consider point-wise properties of the considered predictors. We usually do not know the true distribution on the $L^2[0,T]$ space, from which our data come from. Thus even the existence of the functional expected value (mean) cannot be assumed. Hence, we concentrate our attention on the median-unbiasedness. Let us remind, that an estimate of a one-dimensional parameter is median-unbiased if, for fixed parameter value, the median of the distribution of the estimate is at the parameter value, which simply means that the estimate underestimates just as often as it overestimates (Brown 1947). The classical median-unbiasedness properties have been studied previously (e.g. Pfanzagl 1970, 1979). In the functional setting, we choose the proper functional depth and thus we obtain the median induced by the chosen depth. The functional medians induced by popular depth exist for very wide class of processes (in contrary to the functional mean existence). We conclude, that the functional median obtained with respect to the chosen functional depth is intrinsically a median-unbiased estimator. Moreover, the double median method is not only median-unbiased, but also consistent (for details see Gijbels and Nagy 2015, Nagy et al. 2016 and Kosiorowski et al. 2017c).

Shang and Hyndman (2017) method depends on quite effective but non-robust one-dimensional time series methodology applied to series of principal component scores. It depends also on nonrobust dispersion matrix estimator. The matrix is a kind of a design matrix used to obtain a proper forecasts reconciliation. The robustness of double functional median method to outliers does not heavily depend on the type of functional outliers. It is surprising, because we have expected that it should be different for the functional shape outliers, functional amplitude outliers, and for functional outliers with respect to the covariance structure (e.g. see Arribas-Gil and Romo 2014 and Tarabelloni 2017). After conducting several simulations (see Kosiorowski et al. 2017c) the authors have come to the conclusion, that the double median method is more robust. Shang and Hyndman (2017) state the opposite in their paper, but note, that they considered a Fraiman-Muniz depth, while we have considered MBD, that looks like better designed for the considered empirical example. For a fixed $\alpha$, a volume of the $\alpha$--central region may be treated as a dispersion measure (see Liu et al. 1999), and thus comparing functional boxplots is a relevant way to compare “effectiveness” of the considered methods (see Figures 2, 3, 5, 7). A comparison of functional time series predictor “effectiveness” may be also conducted in terms of speeds of expansion of $\alpha$--central regions treated as functions of $\alpha$ (scale curve, see López-Pintado et al. 2010). This approach is not only nonparametric but it is a moment-free data-analytic method. It imitates the multivariate case and seems
to be the best solution in the functional case as well, because assumptions on the data-generating process are hard to be stated precisely. Remember, that there is no Lebesgue measure analogue in the $L^2[0,T]$ space.

The double median method of forecasting HFTS is faster than Shang and Hyndman (2017) method; moreover, it is less computationally and memory intensive. Precisely, to compare a computational complexity of both methods, we have considered empirical functional time series related to day and night air pollution monitoring in the selected towns. The monitoring was conducted for 181 days. In other words, in the beginning, we considered dataset consisting of six matrices, each of dimension $181 \times 24$. For comparing two forecasting methods, we have considered forecasts obtained basing on moving window of length 10 observations. A time of calculation of the forecasts using Shang and Hyndman (2017) method was ca 13 min 10 sec, whereas using the proposed double median method was ca 2 min 30 sec (we used \textit{DepthProc} R package). In both cases, we used the same software and hardware environment (WIN8, Intel Core I7 Mobile, 16 GB RAM). Note that Shang and Hyndman (2017) and Hyndman \textit{et al.} (2011) indicated the inconveniences of their methods, which are related to an application of the generalized least squares applied to big and sparse design matrices. They stated some methods to bypass the inconveniences, but the remedies are insufficient in big data analysis.

5 Empirical study: The day and night PM 10 air pollution in Silesia region

Air pollution consist of different substances, i.a. sulphur dioxide, nitrogen dioxide, ozone, carbon monoxide, benzene, particulate matter PM2.5 and particulate matter PM10 - all particles of a diameter 10 micrometers or less. Air pollution has a huge negative impact on people’s health.

Air pollution monitoring is conducted in Silesian Province in Poland. Measurement is done at a certain number of stations placed in the Region. The organisation responsible for the monitoring is Wojewódzki Inspektorat Ochrony Środowiska (WIOŚ, Regional Inspectorate of Environmental Protection) in Katowice. The institution possesses 28 measurement stations. We analyze data coming from 5 out of that 28 stations in order to present our method, but the number of stations does not limit our method.

Decision-maker who has at his disposal measurement from certain number of stations is interested in aggregation of the data. Note, that the easiest aggregate is an arithmetic mean or a moving arithmetic mean of measurements done in all the stations. The two aggregates are often used in practice by the local government. However, simplicity seems to be the only advantage of that method (see Section 3 and the paper Kosiorowski \textit{et al.} 2017c). Main goal of the paper is to find the aggregate, to the best fit of the regional policy.
Figure 2: The raw data of 181 curves for the analyzed five stations, which show the PM10 concentration in the atmosphere in [µg/m³]. Above a functional boxplot for all curves (181 × 5 = 905 curves) computed with MBD.
Figure 3: PM10 concentration in the air forecast in $\mu g/m^3$ calculated with moving functional average (moving window equals 10) for five stations. Above a functional boxplot for the average of all averages for five stations computed every day.
Figure 4: PM10 concentration in the air double functional median forecast in µg/m$^3$ calculated with moving functional median (moving window equals 10) for five stations. Above the forecast for the Silesia region calculated with double moving functional median.
5.1 Empirical dataset under study description

Dataset from WIOŚ website [http://powietrze.katowice.wios.gov.pl](http://powietrze.katowice.wios.gov.pl) has been analyzed to illustrate our method.

We have analyzed PM10 concentration in the air for five measurement stations: Gliwice (Gli) with a population of 182,155, Katowice (Kat) with a population of 304,063, Dąbrowa-Górnicza (Dab) with a population of 121,902, Bielsko-Biała (Bie) with a population of 172,407 and Częstochowa (Cze) with a population of 227,184 (population data come from 2015).

First three towns are part of “Upper Silesian Urban Area” (its population is about 3 million). Bielsko-Biała and Częstochowa are the largest towns of Silesian Region that are not part of the “Upper Silesian Urban Area”. Data comes from the period of 181 days from 1 September 2016 to 28 February 2017. We obtain forecasts for each town, but we are rather interested to obtain a forecast for the whole Silesian Region, and we shall keep in mind that emission of pollution and weather conditions (i.e. landform and windrose) are very different in each town. Moreover, some of that factors, i.e. wind, are time variant, so we should treat the observations as functions and treat the trajectories as functional data objects.

It is cumbersome in air pollution context to decide how to compute air pollution on the whole Silesian Region level, as obviously only data from certain stations are available. We decided to compute an aggregate representing air pollution in the Silesian Region to be a weighted average of pollution in each town, where weights are proportional to the town population. This approach is compatible with our assumption that social cost associated with air pollution is linearly proportional to town population. The assumption has been applied in double functional median method and in Shang and Hyndman (2017) method.

Figure 2 presents the raw data of 181 curves for the analyzed five stations, which show the PM10 concentration in the atmosphere in $\mu g/m^3$ on vertical axis. Above there is a functional boxplot for all curves $(181 \times 5 = 905$ curves) computed with MBD.

Figure 3 presents PM10 concentration in the air forecast in $\mu g/m^3$ calculated with moving functional average (moving window equals 10) for five considered stations. Above there is a functional boxplot for the average of all averages for five stations computed every day. The moving average seems to be the easiest rational method of forecasting, which is used by the decision-maker.

Figure 4 presents PM10 concentration in the air double functional median forecast in $\mu g/m^3$ calculated with moving functional median (moving window equals 10) for five stations. Above the forecast for the Silesia region calculated with double moving functional median. The median was calculated with the use of MBD.

Hyndman and Shang (2017), in order to estimate prediction uncertainty, have used maximal entropy bootstrap for time series method proposed by Vinod and de Lacalle (2009), because their method is basing on representation of functional time series as a family of one-dimensional time series of functional principal component scores.

In case of the double functional median method, in order to estimate prediction
uncertainty we have used volumes of $\alpha$–central regions (see functional boxplots) implemented in R-packages fda (Ramsay et al. 2009) and DepthProc (Kosiorowski and Zawadzki 2017). We have also compared quality of our forecast with the forecast of Shang and Hyndman (2017) through the comparison of sum of the differences between the observed curves and of the forecasted curves. We have also compared median absolute deviation (MAD) of the integrated differences between the observed curves and of the forecasted curves. Table 1 contains a MAD comparison of our forecasts with Shang and Hyndman’s (2017). The functional boxplots can be also used to compare the two methods. Figure 5 presents five functional boxplots for the average sum of the differences between the observed curves and the curves forecasted with the double functional median method. Above there is a functional boxplot for the forecasts for Silesia region obtained with the double functional median method (calculated with MBD).
Figure 6: PM10 concentration in the air forecast in µg/m³ calculated with Shang and Hyndman (2017) method (moving window equals 10) for five stations and the Silesia region.

FORECASTS IN WHOLE SILESIA, S&H

67 D. Kosiorowski et al.
CEJME 10: 53-73 (2018)
Figure 6 presents PM10 concentration in the air forecast in $\mu g/m^3$ calculated with Shang and Hyndman (2017) method (moving window equals 10) for five stations and for the Silesia region. Figure 7 presents five boxplots for the average sum of the differences between the observed curves and the curves forecasted with Shang and Hyndman (2017) method. Above a functional boxplot for the forecasts for Silesia region obtained with the Shang and Hyndman (2017) method. Our method seems to be more robust for functional outliers. Compare MAD of the integrated differences between the observed curves and of the curves forecasted with both methods (see Table 1). Compare also boxplots with respect to sizes of the boxes and positions of the medians (see Figures 5 and 7). This result is not very surprising, as Shang and Hyndman (2017) make forecasts basing on nonrobust generalized least squares method.

D. Kosiorowski et al. CEJEME 10: 53-73 (2018)
Table 1: Estimators quality comparison – MAD calculated for the five towns

| Predictor | Biel  | Cze  | Dab  | Gli   | Kat   |
|-----------|-------|------|------|-------|-------|
| Shang&Hyndman | 1699.59 | 1447.81 | 1699.59 | 881.25 | 891.19 |
| Our forecasts    | 265.81  | 377.35 | 265.81  | 526.79  | 328.63 |

5.2 Maximization of social welfare

Generally speaking, it is known that air pollution has a negative impact on human health, however a form of this impact may take many different and very complex forms. Dangerous substances may interact one with another. A nature of impact may depend on age group and time of the day and night.

Our main aim is to maximize a “summarized” utility of a local community over a certain period, that symbolically may be written as (for details see Fleurbaey and Maniquet 2011):

$$U_{Total} = \sum_{i=1}^{365} \int_{[000,2400]} U_i(W_{PM10}(t), C_{PM10reduc}(t)) dt,$$

(2)

where $i$ is a number of a day, $W_{PM10}$ denotes social welfare related to PM10 emission reduction (positive and negative external effects valued in a fixed currency) and $C_{PM10}$ denotes a cost of the PM10 emission reduction valued in the fixed currency.

We assume that

$$W_{PM10} = F(Air_{qual}, ENV_{pol}, INF_{qual}, POP_{param}),$$

and

$$C_{PM10} = G(C_{fixed}, C_{var}, C_{political}, Pred_{qual}).$$

It means that welfare related to PM10 is a function of an air quality (valued basing on evidenced costs of hospitalization due to lung diseases), medical expenses related to allergies ($Air_{qual}$), an user-friendliness of the local environment ($ENV_{pol}$), a quality of the local information system providing information on air quality and health threats ($INF_{qual}$) and finally socio-demographic parameters of the community ($POP_{param}$).

The cost related to the PM10 reduction relates to fixed costs including investments in new technologies ($C_{fixed}$), variable costs involving effects tied with changes of weather causing lower or higher demand on heating energy ($C_{var}$), political cost related to a transformation of popular heating systems basing for example on coal into “clean” systems basing for example on nuclear energy, and costs related to quality of forecasting of air pollution ($Pred_{qual}$).

In this paper we focus our attention on the last quantity measuring quality of forecasting of air pollution in the selected region.
In our opinion it is reasonable to assume that the welfare associated with the pollution in considered town or region is linearly proportional to its population.

6 Conclusions

In the paper we have critically discussed an application of a model for hierarchical functional time series to studies of the day and night air pollution in Silesia region in Poland. We have focused our attention on optimal estimation issues of the model. In this context we have compared the double functional median estimator with the best estimator known from the literature. We have assumed that an aggregate welfare of people living in the Silesia region is a function of among others a quality of the model estimation.

Our considerations clearly show usefulnesses of the HFTS methodology in a context of a local community welfare optimization. Shang and Hyndman (2017) approach provides elegant tools for HFTS modeling and forecasting in case of a relatively rich hierarchical structure and functional data without outliers. It is worth noticing, that the double functional median HFTS predictor performs very well in comparison to Shang and Hyndman (2017) predictor especially in terms of its computational complexity and robustness to functional outliers.

In our current research we concentrate on the optimization issues related to specific forms of the formula (2) defining the welfare of a certain local community.

Acknowledgements

The authors would like to thank two anonymous referees for their valuable suggestions and helpful comments.

JPR and DM’s research has been partially supported by the AGH UST local grant no. 11.11.420.004 and DK’s research by the grant awarded to the Faculty of Management of CUE for preserving scientific resources for 2017 and 2018.

References

[1] Arribas-Gil A., Romo J. (2014), Shape outlier detection and visualization for functional data: the outliergram, *Biostatistics* 15(4), 603–619.

[2] Brown G.W., (1947), On Small-Sample Estimation, *Annals of Mathematical Statistics* 18(4), 582–585.

[3] Ferraty F., and Vieu P. (2006), *Nonparametric Functional Data Analysis: Theory and Practice*, Springer.
Forecasting of a Hierarchical Functional Time Series . . .

[4] Fleurbaey M., Maniquet F. (2011), *A Theory of Fairness and Social Welfare*, Cambridge University Press.

[5] Fliedner G. (2001), Hierarchical forecasting: issues and use guidelines, *Industrial Management and Data Systems* 101(1), 5–12.

[6] Gijbels I., Nagy S. (2015), Consistency of non-integrated depths for functional data, *Journal of Multivariate Analysis* 140, 259–282.

[7] Górecki T., Krzyśko M.K., Waszak Ł., Wołyński W. (2016), Selected statistical methods of data analysis for multivariate functional data, *Statistical Papers*, DOI 10.1007/s00362-016-0757-8.

[8] Horváth L., Kokoszka P. (2012), *Inference for Functional Data with Applications*, Springer, New York.

[9] Hyndman R.J., Ahmed R.A., Athanasopoulos G., Shang H.L. (2011), Optimal combination forecasts for hierarchical time series, *Computational Statistics & Data Analysis* 55(9), 2579–2589.

[10] Kahn K.B. (1998), Revisiting top-down versus bottom-up forecasting, *The Journal of Business Forecasting Methods & Systems* 17(2), 14–19.

[11] Kohn R. (1982), When is an aggregate of a time series efficiently forecast by its past, *Journal of Econometrics* 18(3), 337–349.

[12] Kosiorowski D. (2014), Functional regression in short-term prediction of economic time series, *Statistics in Transition – new series* 15(4), 611–626.

[13] Kosiorowski D. (2016), Dilemmas of robust analysis of economic data streams, *Journal of Mathematical Sciences* 1(2), 59–72.

[14] Kosiorowski D., Zawadzki Z. (2017), DepthProc: An R Package for Robust Exploration of Multidimensional Economic Phenomena, available at: arXiv:1408.4542v9.

[15] Kosiorowski D., Rydlewski J.P., Snarska M. (2017a), Detecting a structural change in functional time series using local Wilcoxon statistic, *Statistical Papers*, DOI 10.1007/s00362-017-0891-y.

[16] Kosiorowski D., Mielczarek D., Rydlewski J.P., Snarska M. (2017b), Generalized exponential smoothing in prediction of hierarchical time series, available at: arXiv:1612.02195v2.

[17] Kosiorowski D., Mielczarek D., Rydlewski J. P. (2017c), Double functional median in robust prediction of hierarchical functional time series – an application to forecasting of the Internet service users behaviour, available at: arXiv:1710.02669v1.

71 D. Kosiorowski et al.
CEJEME 10: 53-73 (2018)
Daniel Kosiorowski, Dominik Mielczarek, Jerzy P. Rydlewski

[18] Kosiorowski D., Rydlewski J.P., Zawadzki Z. (2017d), Functional outliers detection by the example of air quality monitoring, *Statistical Review* (in Polish, forthcoming).

[19] Liu R., Parelius J.M., Singh K. (1999), Multivariate analysis by data depth: descriptive statistics, graphics and inference (with discussion), *Annals of Statistics* 27, 783–858.

[20] López-Pintado S., Romo J. (2007), Depth-based inference for functional data, *Computational Statistics & Data Analysis* 51(10), 4957–4968.

[21] López-Pintado S., Romo J. (2009), On the concept of depth for functional data, *Journal of the American Statistical Association* 104, 718–734.

[22] López-Pintado S., Romo J., Torrente A. (2010), Robust depth-based tools for the analysis of gene expression data, *Biostatistics* 11(2), 254–264.

[23] Nagy S., Gijbels I., Omelka M., Hlubinka D. (2016), Integrated depth for functional data: Statistical properties and consistency, *ESIAM Probability and Statistics* 20, 95–130.

[24] Nieto-Reyes A., Battey H. (2016), A Topologically Valid Definition of Depth for Functional Data, *Statistical Science* 31(1), 61–79.

[25] Pfanzagl J. (1970), On the Asymptotic Efficiency of Median Unbiased Estimates, *The Annals of Mathematical Statistics* 41(5), 1500–1509.

[26] Pfanzagl J. (1979), On optimal median unbiased estimators in the presence of nuisance parameters, *The Annals of Statistics* 7(1), 187–193.

[27] Ramsay J.O., Hooker G., Graves S. (2009), *Functional Data Analysis with R and Matlab*, Springer.

[28] Shang H.L., Hyndman R.J. (2017), Grouped functional time series forecasting: an application to age-specific mortality rates, *Journal of Computational and Graphical Statistics* 26(2), 330–343.

[29] Shlifer E., Wolff R.W. (1979), Aggregation and proration in forecasting, *Management Science* 25(6), 594–603.

[30] Sun Y., Genton M.G. (2011), Functional Boxplots, *Journal of Computational and Graphical Statistics* 20(2), 316–334.

[31] Tarabelloni N. (2017), Robust Statistical Methods in Functional Data Analysis, Doctoral thesis and R package *roahd*, Politecnico di Milano.

[32] Vinod H.D., Lopez de Lacalle J. (2009), Maximum entropy bootstrap for time series: the meboot R package, *Journal of Statistical Software* 29(5), 1–19.
[33] Weale M. (1988), The reconciliation of values, volumes and prices in the national accounts, *Journal of the Royal Statistical Society A* 151(1), 211–221.

[34] Zuo Y., Serfling R. (2000), General notions of statistical depth function, *Annals of Statistics* 28(2), 461–482.

[35] [http://powietrze.katowice.wios.gov.pl](http://powietrze.katowice.wios.gov.pl) (access date: 24th March 2017).

WIOS underlines that the data is gathered automatically and might be unverified.