Associations between three common single nucleotide polymorphisms (rs266729, rs2241766, and rs1501299) of ADIPOQ and cardiovascular disease: a meta-analysis

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Abstract

Background: Inconsistencies have existed in research findings on the association between cardiovascular disease (CVD) and single nucleotide polymorphisms (SNPs) of ADIPOQ, triggering this up-to-date meta-analysis.

Methods: We searched for relevant studies in PubMed, EMBASE, Cochrane Library, CNKI, CBM, VIP, and WanFang databases up to 1st July 2017. We included 19,106 cases and 31,629 controls from 65 published articles in this meta-analysis. STATA 12.0 software was used for all statistical analyses.

Results: Our results showed that rs266729 polymorphism was associated with the increased risk of CVD in dominant model or in heterozygote model; rs2241766 polymorphism was associated with the increased risk of CVD in the genetic models (allelic, dominant, recessive, heterozygote, and homozygote). In subgroup analysis, significant associations were found in different subgroups with the three SNPs. Meta-regression and subgroup analysis showed that heterogeneity might be explained by other confounding factors. Sensitivity analysis revealed that the results of our meta-analysis were stable and robust. In addition, the results of trial sequential analysis showed that evidences of our results are sufficient to reach concrete conclusions.

Conclusions: In conclusion, our meta-analysis found significant increased CVD risk is associated with rs266729 and rs2241766, but not associated with rs1501299.

Keywords: ADIPOQ, Single nucleotide polymorphisms, Cardiovascular disease, Association, Meta-analysis

Background

Cardiovascular disease (CVD) is the primary cause of death worldwide, leading to 32% of all deaths worldwide in 2013 [1]. Epidemiological and biological evidences demonstrate that multiple environmental and genetic factors are implicated in CVD, although the etiology of CVD has not been fully elucidated [2–5]. Identifying CVD-relative risk factors is critical in control of the development and progress of CVD.

Adiponectin is involved in CVD: low levels of adiponectin (hypoadipoectinemia) positively correlate with the risk of CVD, and higher levels of adiponectin protect against this disease [6–11]. Adiponectin is synthesized and secreted by adipose tissue [12], osteoblasts [13], skeletal muscle [14], and cardiomyocytes [15]. This protein, as one of the most abundant adipocytokines in blood, has anti-atherogenic, cardioprotective, anti-inflammatory, and antithrombotic properties [16–20].

Adiponectin is encoded by ADIPOQ which is located in chromosome 3q27 [21], and adiponectin levels are influenced by single-nucleotide polymorphisms (SNPs) in ADIPOQ [22]. SNPs in ADIPOQ have been found to be associated with CVD [23, 24], diabetes [25, 26], stroke [27, 28], myocardial infarction [29, 30], cancer [31, 32],
kidney disease [33, 34], and even gynecological conditions [35, 36]. Previous studies have shown the association between SNPs in ADIPOQ (rs3774261, rs1063537, rs2082940, rs2241766, rs266729, and rs1501299) and CVD/subclinical CVD [30, 37, 38]. The three common SNPs of ADIPOQ (rs266729, rs2241766, and rs1501299) were most widely studied. However, findings from previous studies on the three SNPs in relation to CVD risk are inconsistent and inconclusive.

For rs266729 (11,377 C/G) in ADIPOQ, Du et al. [39] and Zhang et al. [40] found that the SNP is associated with CVD risk; Stenvinkel et al. [41] revealed that rs266729 is associated with the decreased risk of CVD; Zhang et al. [40], Cheong et al. [27], and Chioldini et al. [29] found that there is no significant association between rs266729 and CVD. For rs2241766 (+45 T/G), Pischon et al. [42] and Jung et al. [43] identified no association between rs2241766 and the risk of coronary artery disease (CAD) in patients with type 2 diabetic mellitus (T2DM); Du et al. [39], Oliveira et al. [44], and Mofarrah et al. [45] found that there is a significant association between rs2241766 polymorphism and CAD risk; Chang et al. [46] revealed that rs2241766 is associated with the decreased risk of CVD. Moreover, for rs1501299 (+276 G/T), Bacci et al. [47] and Esteghamati et al. [48] revealed that rs1501299 is associated with the decreased risk of CAD; Mohammadzadeh et al. [38], however, reported that there is an association between rs1501299 and CAD risk; Foucan et al. [49] found that there is no significant association between rs1501299 and CAD in patients with T2DM. Thus, those results are inconsistent.

Meta-analysis performed by Zhang et al. in 2012 revealed that associations between the SNPs (rs2241766, rs1501299, and rs266729) in ADIPOQ and CVD were significant but weak [50]. Since that data, several more studies have emerged to investigate the association between SNPs in ADIPOQ and susceptibility to CVD [37, 38, 45]. In this study, we further collected references and updated meta-analysis of association between SNPs (rs2241766, rs1501299, and rs266729) in ADIPOQ and CVD in order to get a more precise and reliable assessment of the association.

Methods
Search strategy
We performed an extensive literature search in PubMed, EMBASE, Cochrane Library, CNKI, CBM, VIP, and WanFang databases for published articles on the association between ADIPOQ polymorphisms and CVD risk up to July 1st, 2017. The literature search was done without any language or population restrictions imposed. During the literature search, we used various combinations of keywords, such as ‘coronary heart disease (CHD)’ or ‘coronary artery disease’ or ‘cardiovascular disease’ or ‘ischemic heart disease’ or ‘angina’ or ‘myocardial infarction (MI)’ or ‘stroke’ or ‘atherosclerosis’ or ‘arteriosclerosis’ or ‘coronary stenosis’ combined with ‘ADIPOQ’ or ‘APM1’ or ‘ACDC’ or ‘adiponectin gene’ and ‘polymorphisms’ or ‘variants’ or ‘variations’. Joseph Sam Kanu and Shuang Qiu independently performed the literature search for potential articles included in this meta-analysis. All articles retrieved were first organized in reference manager software (Endnote 6).

Inclusion and exclusion criteria
A study included in this meta-analysis was based on the following criteria: 1) the study has sufficient data to allow association between CVD risk and ADIPOQ SNP to be assessed; 2) the study included original data (independence among studies); 3) evaluation of the ADIPOQ polymorphisms (rs266729, rs2241766, and rs1501299) and CVD risk; 4) the language of the study was English or Chinese; and 5) observed genotype frequencies in controls must be consistent with Hardy–Weinberg equilibrium (HWE). We excluded a study based on: 1) the study contained overlapping data; 2) the study with missing information (particularly genotype distributions), after corresponding author, who was contacted by us with email, failed to provide the required information; and 3) genome scans investigating linkages with no detailed genotype distributions between cases and controls. Where there was a disagreement on the selection of a study, the issue was resolved by discussion or consensus with the third investigator (Ri Li). For articles with missing data, we emailed the corresponding authors for the required data.

Assessment of study quality
We used the NATURE-published guidelines proposed by the NCI-NHGRI Working Group on Replication in Association Studies for assessing the quality of each study included in this meta-analysis [51]. These guidelines have a checklist of 53 conditions for authors, journal editors, and referees to interpret data and results of genome-wide or other genotype–phenotype association studies clearly and unambiguously. We used the first set of 34 conditions in assessing the quality of each study. We allocated a score of 1 point for each condition a study met, and no point (0 score) if the condition or requirement is lacking. Each study was given a total Quality Score – the sum of all points each study obtained. Study quality assessment was independently carried out by Joseph Sam Kanu and Shuang Qiu.

Data extraction
Joseph Sam Kanu and Shuang Qiu extracted data from each study independently. We summarized the information extracted from each article in Table 1. The characteristics
of articles included first author, year of publication, country in which the study was done, study population (ethnicity), numbers of cases and controls, genotyping method, SNPs investigated, genotype frequency of cases and controls, and outcome (Table 1; Additional file 1: Tables S1, S2, and S3).

**Statistical analysis**

HWE was evaluated for each study using Goodness of fit Chi-square test in control groups, and $P < 0.05$ was considered as a significant deviation from HWE. The strength of association between the three ADIPOQ polymorphisms and CVD susceptibility was assessed using odds ratios ($OR$) and 95% confidence intervals (95% $CI$). The associations were measured based on five different genetic models: allelic model ($rs266729$: G versus C; $rs2241766$: G versus T; $rs1501299$: T versus G), dominant model ($rs266729$: GG versus GC + CC; $rs2241766$: GG + GT versus TT; $rs1501299$: TT + TG versus GG), recessive model ($rs266729$: GG versus GC + CC; $rs2241766$: GG + GT versus TT; $rs1501299$: TT + TG versus GG), heterozygote model ($rs266729$: GC versus CC; $rs2241766$: GT versus TT; $rs1501299$: TG versus GG), and homozygote model ($rs266729$: GG versus GC + CC; $rs2241766$: GG versus GT + GG; $rs1501299$: TT versus TG + GG), and homozygote model ($rs266729$: GG versus GC + CC; $rs2241766$: GG versus GT + GG; $rs1501299$: TT versus TG + GG). Heterogeneity were evaluated by the Chi-square test based Q-statistic, and quantified by $I^2$-statistic [52]. If there was no substantial statistical heterogeneity ($P > 0.10$, $I^2 \leq 50\%$), data were pooled by fixed-effect model (Mantel and Haenszel methods); otherwise, the heterogeneity was evaluated by random-effect model (DerSimonian and Laird methods). Meta-regression analysis was performed to detect main sources of heterogeneity. In addition, subgroup analyses were stratified by population (European, East Asian, West Asian, and African), genotyping method (PCR-RFLP, TaqMan, and Others), sample size (<1000 and ≥1000), and quality score (<10 and ≥10). Sensitivity analysis was performed to examine stability of our results by omitting each study in each turn. Publication bias was measured by funnel plots [53], and quantified by the Begg’s and Egger’s tests [54] ($P < 0.05$ considered statistically significant publication bias). STATA 12.0 software (StataCorp. 2011. Stata Statistical Software: Release 12. College Station, TX: StataCorp LP) was used for all statistical analyses. $P$-value <0.05 was considered statistically significant, except where other-wise specified. A separate analysis was performed for each SNPs included in the meta-analysis.

**Trial sequential analysis (TSA)**

Traditional meta-analysis may result in type I and type II errors owing to dispersed data and repeated significance testing [55, 56]. To reduce the risk of type I error, TSA was used to estimate required information size (RIS) and confirm statistical reliability with an adjusted threshold for statistical significance [57]. In present meta-analysis, we used trial sequential analysis software (TSA, version 0.9; Copenhagen Trial Unit, Copenhagen, Denmark, 2011) by setting an overall type I error of 5%, a statistical test power of 80%, and a relative risk reduction of 20% [58, 59]. If the Z-curve crosses trial sequential monitoring boundary or RIS has been reached, a sufficient level of evidence has been reached and further studies are unnecessary; otherwise, additional studies are needed to reach a sufficient conclusion.

**Results**

**Overall results**

This meta-analysis included 68 studies from 65 articles after literature search and critical screening, as described in methods (Fig. 1). Meta-analysis of the $rs266729$ (−11,377 C > G), $rs2241766$ (+45 T > G), and $rs1501299$ (+276 G > T) variants included 29, 40, and 44 studies, respectively. We summarize the characteristics of each primary study in Table 1. Detailed characteristics of those studies are further presented in Additional file 1: Tables S1, S2, and S3. Overall, this meta-analysis included a total of 50,735 subjects (19,106 cases and 31,629 controls).

**Meta-analysis results**

**Association between $rs266729$ (−11,377 C > G) polymorphism and CVD**

The meta-analysis of the association between $rs266729$ (−11,377 C > G) polymorphism and CVD included 29 studies with 29,021 subjects (10,506 cases and 18,515 controls). Significant heterogeneity among studies was observed ($P_h < 0.10$ or $I^2 \geq 50\%$). Thus, we selected random-effect model, and found that $rs266729$ polymorphism was associated with the increased risk of CVD in dominant model ($GC + GC$ VS CC: $OR = 1.129$, 95% $CI = 1.028–1.239$, $P = 0.011$) and in heterozygote model ($GC$ VS CC: $OR = 1.141$, 95% $CI = 1.041–1.250$, $P = 0.005$) (Table 2, Fig. 2).

Based on population, genotyping method, sample size, and quality score, we performed subgroup analyses. On the basis of population, $rs266729$ polymorphism was associated with the increased risk of CVD under dominant model ($GC + GC$ VS CC: $OR = 1.198$, 95% $CI = 1.066–1.427$, $P = 0.043$) and under heterozygote model ($GC$ VS CC: $OR = 1.184$, 95% $CI = 1.002–1.398$, $P = 0.048$) in East Asian. On the basis of genotyping methods, a significant risk association between $rs266729$ polymorphism and CVD was found when genotyping was performed using PCR-RFLP method under dominant model ($GC + GC$ VS CC: $OR = 1.276$, 95% $CI = 1.014–1.607$, $P = 0.038$) and under heterozygote model ($GC$ VS CC: $OR = 1.282$, 95% $CI = 1.032–1.592$, $P = 0.025$). On the basis of sample
| Study                  | ID | Year | Country      | Population   | Outcome   | Sample size | Genotyping Method | Quality Score |
|-----------------------|----|------|--------------|--------------|-----------|-------------|-------------------|---------------|
| Lacquemant Swiss      | 1  | 2004 | Switzerland  | European     | CAD       | 107         | Other             | 9             |
| Lacquemant French     | 2  | 2004 | France       | European     | CAD       | 55          | Other             | 9             |
| Bacci                 | 3  | 2004 | Italy        | European     | CAD       | 142         | Other             | 8             |
| Ohashi                | 4  | 2004 | Japan        | East Asian   | CAD       | 383         | TaqMan            | 7             |
| Stenvinkel            | 5  | 2004 | America      | European     | CVD       | 63          | Other             | 6             |
| Filippi               | 6  | 2005 | Italy        | European     | CAD       | 980         | Other             | 9             |
| Ru Y                  | 7  | 2005 | China        | East Asian   | CHD       | 131         | TaqMan            | 6             |
| Qi H                  | 8  | 2005 | America      | European     | CVD       | 239         | TaqMan            | 10            |
| Qi Q                  | 9  | 2005 | America      | European     | CVD       | 285         | TaqMan            | 10            |
| Wang JN               | 10 | 2006 | China        | East Asian   | CHD       | 120         | PCR-RFLP          | 7             |
| Hegener               | 11 | 2006 | America      | European     | MI        | 341         | TaqMan            | 11            |
| Hegener               | 12 | 2006 | America      | European     | Stroke    | 259         | TaqMan            | 11            |
| Jung                  | 13 | 2006 | Korea        | East Asian   | CAD       | 88          | TaqMan            | 8             |
| Gable 1               | 14 | 2007 | UK           | European     | CVD       | 266         | PCR-RFLP          | 11            |
| Gable 2               | 15 | 2007 | UK           | European     | MI        | 530         | PCR-RFLP          | 12            |
| Pischon               | 16 | 2007 | America      | European     | CHD       | 1,036       | TaqMan            | 11            |
| Lu F                  | 17 | 2007 | China        | East Asian   | CHD       | 135         | PCR-RFLP          | 7             |
| Hoeft                 | 18 | 2007 | Austria      | European     | CHD       | 277         | TaqMan            | 7             |
| Yamada                | 19 | 2008 | Japan        | East Asian   | ACI       | 313         | Other             | 9             |
| Ogun                  | 20 | 2009 | Japan        | East Asian   | MI        | 773         | Other             | 10            |
| Chang                 | 21 | 2009 | China        | East Asian   | CAD       | 600         | PCR-RFLP          | 9             |
| Zhang XL              | 22 | 2009 | China        | East Asian   | CAD       | 205         | PCR-RFLP          | 8             |
| Zhong C               | 23 | 2010 | China        | East Asian   | CAD       | 198         | TaqMan            | 10            |
| Foucan                | 24 | 2010 | France       | African      | CAD       | 57          | TaqMan            | 7             |
| Xu L                  | 25 | 2010 | China        | East Asian   | CHD       | 153         | PCR-RFLP          | 8             |
| Chiodini              | 26 | 2010 | Italy        | European     | MI        | 503         | TaqMan            | 10            |
| Persson               | 27 | 2010 | Sweden       | European     | MI        | 244         | TaqMan            | 9             |
| Chen XL               | 28 | 2010 | China        | East Asian   | Stroke    | 357         | TaqMan            | 8             |
| Luo SX                | 29 | 2010 | China        | East Asian   | CHD       | 221         | PCR-RFLP          | 8             |
| Caterina              | 30 | 2011 | Italy        | European     | MI        | 1,864       | Other             | 13            |
| Al-Daghri             | 31 | 2011 | Saudi A.     | West Asian   | CAD       | 123         | PCR-RFLP          | 8             |
| Prior                 | 32 | 2011 | UK           | European     | CHD       | 85          | PCR-RFLP          | 7             |
| Leu                   | 33 | 2011 | China        | East Asian   | Stroke    | 80          | Other             | 10            |
| Liu F                 | 34 | 2011 | China        | East Asian   | Stroke    | 302         | PCR-RFLP          | 9             |
| Rodriguez             | 35 | 2011 | Spain        | European     | CVD       | 119         | TaqMan            | 9             |
| Chen F                | 36 | 2011 | China        | East Asian   | CHD       | 93          | PCR-RFLP          | 8             |
| Maimaitiyiming        | 37 | 2011 | China        | East Asian   | CHD       | 196         | PCR-RFLP          | 8             |
| Hu HF                 | 38 | 2011 | China        | East Asian   | CHD       | 150         | PCR-RFLP          | 8             |
| Zhang YM              | 39 | 2011 | China        | East Asian   | CHD       | 149         | PCR-RFLP          | 8             |
| Zhou NN               | 40 | 2011 | China        | East Asian   | CAD       | 358         | PCR-RFLP          | 8             |
| Sabouri               | 41 | 2011 | UK           | European     | CAD       | 329         | PCR-RFLP          | 8             |
| Boumaiza              | 42 | 2011 | Tunisia      | African      | CAD       | 212         | PCR-RFLP          | 10            |
| Chengang              | 43 | 2012 | China        | East Asian   | CAD       | 267         | PCR-RFLP          | 8             |
size or quality score, we found that rs266729 polymorphism was associated with the increased risk of CVD under allelic, dominant, and heterozygote models (all OR > 1 and P < 0.05), after pooled the ORs by the subgroups of sample size ≥ 1000 or quality score ≤ 10 (Table 2).

**Association between rs2241766 (+45 T > G) polymorphism and CVD**

The meta-analysis of the association between rs2241766 (+45 T > G) polymorphism and CVD included 40 studies with 25,548 subjects (10,746 cases and 14,802 controls). Using inverse-variance weighted random effect model (I^2 ≥ 50%), we found that rs2241766 polymorphism was associated with the increased risk of CVD in the five genetic models (allelic, dominant, recessive, heterozygote, and homozygote) (all OR > 1 and P < 0.05) (Table 3, Fig. 3).

Subgroup analyses were stratified by population, genotyping method, sample size, and quality score. Firstly, on the basis of population, rs2241766 polymorphism was associated with the increased risk of CVD under the five dominant models in East Asian and under allelic, recessive, and homozygote models in West Asian (all OR > 1 and P < 0.05). Secondly, on the basis of genotyping method, the results that genotyping was done by PCR-RFLP or other methods showed that rs2241766 polymorphism was associated with the increased risk of CVD under five genetic models (all OR > 1 and P < 0.05). Thirdly, on the basis of sample size, rs2241766 polymorphism was associated with the increased risk of CVD under the five genetic models in the subgroup of sample size ≤ 1000 (all OR > 1 and P < 0.05), but was associated with the decreased risk of CVD in the subgroup of sample size ≥ 1000 under recessive model (GG VS...
GT + TT: OR = 0.696, 95% CI = 0.539–0.885, \( P = 0.003 \) and under homozygote model (GG VS TT: OR = 0.669, 95% CI = 0.519–0.862, \( P = 0.002 \)). Finally, on the basis of quality score, when we pooled the O Rs by the subgroups of quality score \( \leq 10 \), we found that rs2241766 polymorphism was associated with the increased risk of CVD under the five genetic models (all OR > 1 and \( P < 0.05 \)) (Table 3).

**Association between rs1501299 (+276 G > T) polymorphism and CVD**

The meta-analysis of the association between rs1501299 (+276 G > T) polymorphism and CVD included 44 studies with 37,371 subjects (12,852 cases and 24,519 controls). Using the inverse-variance weighted random effect model (\( P_h < 0.10 \) or \( I^2 \geq 50\% \)), we found that there was no association between rs1501299 polymorphism and CVD in the five genetic models (all \( P > 0.05 \)) (Table 4). In the subgroup analysis, no significant association was found between rs1501299 polymorphism and CVD risk under the five genetic models in any subgroup (all \( P > 0.05 \)) (Table 4).

**Heterogeneity analysis**

In this meta-analysis, meta-regression was used to investigate the source of heterogeneity by year, population, genotyping method, sample size, and quality score. We found that sample size (allelic model: \( P = 0.019 \); dominant model: \( P = 0.032 \); recessive model: \( P < 0.001 \); and homozygote model: \( P < 0.001 \)) and quality score (allelic model: \( P = 0.035 \); dominant model: \( P = 0.032 \); recessive model: \( P < 0.001 \); and homozygote model: \( P < 0.001 \)) contributed to the observed heterogeneity across all the
| Categories | n  | Sample size | G VS C | OR (95% CI) | P   | G + GC VS CC | OR (95% CI) | P   | GG VS CC | OR (95% CI) | P   |
|------------|----|-------------|--------|-------------|-----|--------------|-------------|-----|---------|-------------|-----|
| Overall    | 29 | 10,506/18,515 | 1.079 | (1.000, 1.165) | 0.051 | 65.8/0.000 | 1.129 | (1.028, 1.239) | 0.011 | 645/0.000 | 0.899 | (0.838, 1.168) |
| Population |    |             |        |             |     |              |             |     |         |             |     |
| European   | 17 | 6,355/11,666 | 1.022 | (0.948, 1.102) | 0.564 | 376/0.060   | 1.071 | (0.974, 1.178) | 0.158 | 408/0.041 | 0.879 | (0.714, 1.082) |
| East Asian  | 12 | 4,151/6,849 | 1.154 | (1.000, 1.332) | 0.051 | 768/0.000   | 1.198 | (1.006, 1.427) | 0.043 | 757/0.000 | 1.149 | (0.887, 1.487) |
| Genotyping |    |             |        |             |     |              |             |     |         |             |     |
| PCR-RFLP   | 8  | 2,382/4,976 | 1.186 | (0.978, 1.438) | 0.083 | 775/0.000   | 1.276 | (1.014, 1.607) | 0.038 | 754/0.000 | 1.162 | (0.813, 1.661) |
| TaqMan     | 12 | 3,910/6,312 | 1.031 | (0.935, 1.135) | 0.544 | 453/0.044   | 1.054 | (0.948, 1.173) | 0.331 | 307/0.146 | 0.951 | (0.730, 1.259) |
| Others     | 9  | 4,214/7,227 | 1.045 | (0.921, 1.186) | 0.493 | 636/0.005   | 1.095 | (0.926, 1.296) | 0.289 | 687/0.001 | 0.923 | (0.711, 1.197) |
| Sample size|    |             |        |             |     |              |             |     |         |             |     |
| < 1000     | 21 | 5,048/6,708 | 1.065 | (0.952, 1.192) | 0.270 | 679/0.000   | 1.114 | (0.973, 1.276) | 0.119 | 659/0.000 | 0.955 | (0.744, 1.228) |
| ≥ 1000     | 8  | 5,458/11,807 | 1.108 | (1.004, 1.222) | 0.042 | 634/0.008   | 1.162 | (1.026, 1.315) | 0.018 | 646/0.006 | 1.017 | (0.835, 1.240) |
| Quality score|   |          |        |             |     |              |             |     |         |             |     |
| < 10       | 16 | 3,489/5,128 | 1.152 | (1.007, 1.318) | 0.040 | 680/0.000   | 1.211 | (1.032, 1.420) | 0.019 | 646/0.000 | 1.147 | (0.861, 1.528) |
| ≥ 10       | 13 | 7,017/13,387 | 1.019 | (0.940, 1.105) | 0.064 | 555/0.000   | 1.062 | (0.954, 1.182) | 0.271 | 613/0.002 | 0.883 | (0.744, 1.046) |

n study numbers. Bold values represent statistically significant findings.
studies of the association between rs2241766 polymorphisms and CVD risk. However, in the meta-analysis of the associations between rs266729/rs1501299 polymorphisms and CVD risk, we did not identify the source of heterogeneity (all $P > 0.05$) (Additional file 2: Table S4).

Publication bias and sensitivity analysis

Publication bias was measured by funnel plots and quantified by Begg’s and Egger’s tests. No publication bias was found among the studies regarding the association between rs266729 polymorphisms and CVD risk (all $P > 0.05$). Publication biases were found in analyses of the associations between rs2241766 polymorphisms and CVD risk (allelic model: $P_{Egger} = 0.001$, $P_{Begg} = 0.031$; dominant model: $P_{Egger} = 0.001$, $P_{Begg} = 0.003$; and heterozygote mode: $P_{Egger} = 0.003$, $P_{Begg} = 0.003$), and between rs1501299 polymorphisms and CVD risk (recessive model: $P_{Egger} = 0.031$, $P_{Begg} = 0.035$) (Table 5 and Additional file 3: Figures S1, S2, and S3). Sensitivity analyses showed that this meta-analysis was relatively stable and credible (Figs. 4, 5, and 6).

TSA

In the TSA of rs266729 and CVD, the Z-curve crossed trial sequential monitoring boundary and the sample size reached RIS in dominant and heterozygote models (Fig. 7). In allelic, recessive, and homozygote models, the sample size also reached RIS, although the Z-curve did not cross trial sequential monitoring boundary (Fig. 7). In the TSA of rs2241766/rs1501299 and CVD, the sample size reached RIS in the five genetic models (Figs. 8 and 9). Thus, concrete conclusions were reached and further studies were not required.

Discussion

In this meta-analysis, we collected up-to-date information (July 1st, 2017) to investigate the association between ADIPOQ SNPs and the risk of CVD. Our results demonstrate that rs266729 and rs2241766 variants of ADIPOQ are associated with the increased risk of CVD, but rs1501299 is not associated with CVD risk.

In view of the association between rs266729 and CVD risk, Yang et al. (2012) [60], Zhou et al. (2012) [61], and Zhang et al. (2012) [50] performed meta-analyses. Yang et al. reported that rs266729 is associated with the increased risk of CAD in allelic and dominant models [60]. Zhou et al. found the same association in overall population, Europeans, and East Asian in allelic, dominant and heterozygote models [61]. Zhang et al. also revealed that rs266729 is associated with the increased risk of CAD in overall population and East Asian in allelic model [50]. Our results further identified that rs266729 is associated with the increased risk of CVD in overall population and East Asian in dominant and heterozygote models. In addition, our results revealed that the significant association in studies on the basis of PCR-RFLP method, indicating that different genotyping method may result in different statistical results.

The association between rs2241766 and CVD risk also has been the subject of meta-analysis [60–64]. These studies are inconsistent. Yang et al. found no significant association between rs2241766 and CAD risk [60]. Zhang et al. found no overall significant risk association between CHD and rs2241766 in Han Chinese population [62]. Zhou et al. reported that rs2241766 is associated with the decreased risk of CVD in recessive and homozygote models, and the decreased risk of CVD in East Asian in
Table 3 Overall and subgroup meta-analysis of the association between ADIPOQ rs2241766, +45 T > G polymorphisms and CVD

| Categories | n   | Sample size | G VS T | G + GT VS TT | GG VS GT + T | GT VS TT | GG VS TT |
|------------|-----|-------------|--------|--------------|-------------|----------|---------|
|            |     |             | OR (95% CI) | P    | OR (95% CI) | P    | OR (95% CI) | P    | OR (95% CI) | P    | OR (95% CI) | P    |
| Overall    | 40  | 10,746/14,802 | 1.216 (1.102, 1.343) | < 0.001 | 1.229 (1.103, 1.369) | < 0.001 | 1.286 (1.001, 1.560) | 0.011 | 49.7/0.000 | 1.172 (1.063, 1.292) | 0.001 | 53.3/0.000 | 1.361 (1.095, 1.690) | 0.005 | 57.7/0.000 |
| Population |     |             |        |              |             |         |                      |      |              |       |             |       |
| European   | 12  | 4,452/7,255  | 1.067 (0.918, 1.242) | 0.398 | 60.4/0.003 | 1.105 (0.937, 1.303) | 0.238 | 583/0.006 | 0.779 (0.576, 1.059) | 0.106 | 0.00/668 | 1.123 (0.956, 1.319) | 0.157 | 53.0/0.001 | 0.792 (0.584, 1.073) | 0.132 | 0.0/0585 |
| East Asian | 20  | 5,305/6,505  | 1.194 (1.057, 1.348) | 0.004 | 70.5/0.000 | 1.225 (1.057, 1.420) | 0.007 | 673/0.000 | 1.315 (1.068, 1.618) | 0.010 | 43.1/0.024 | 1.180 (1.029, 1.353) | 0.018 | 58.1/0.001 | 1.431 (1.112, 1.842) | 0.005 | 58.6/0.001 |
| West Asian | 5   | 660/719     | 1.550 (1.002, 2.396) | 0.049 | 80.8/0.000 | 1.392 (0.883, 2.170) | 0.145 | 71.3/0.007 | 2.715 (1.452, 5.079) | 0.002 | 50.2/0.091 | 1.099 (0.779, 1.549) | 0.591 | 43.2/0.134 | 2.767 (1.347, 5.683) | 0.006 | 59.2/0.004 |
| African    | 3   | 329/323     | 2.200 (0.890, 5.437) | 0.088 | 75.3/0.017 | 2.148 (0.952, 4.844) | 0.066 | 65.2/0.056 | 20.10 (0.251, 16.080) | 0.511 | 50.9/0.154 | 1.919 (0.998, 3.688) | 0.051 | 45.1/0.162 | 2.295 (0.250, 21.058) | 0.463 | 55.2/0.135 |
| Genotyping |     |             |        |              |             |         |                      |      |              |       |             |       |
| PCR-RFLP   | 20  | 4,814/6,319  | 1.242 (1.055, 1.462) | 0.009 | 77.1/0.000 | 1.279 (1.057, 1.548) | 0.012 | 744/0.000 | 1.335 (1.034, 1.722) | 0.027 | 405/0.035 | 1.221 (1.023, 1.458) | 0.027 | 67.0/0.000 | 1.442 (1.094, 1.975) | 0.022 | 57.1/0.001 |
| TaqMan     | 7   | 2,616/3,715  | 1.087 (0.895, 1.320) | 0.400 | 64.8/0.000 | 1.118 (0.920, 1.357) | 0.262 | 539/0.048 | 0.872 (0.513, 1.482) | 0.614 | 569/0.041 | 1.123 (0.956, 1.329) | 0.172 | 34.9/0.162 | 0.896 (0.506, 1.588) | 0.708 | 62.3/0.021 |
| Other      | 13  | 3,316/4,768  | 1.263 (1.075, 1.485) | 0.005 | 68.7/0.000 | 1.238 (1.066, 1.452) | 0.009 | 514/0.016 | 1.453 (1.021, 2.066) | 0.038 | 567/0.006 | 1.150 (1.004, 1.317) | 0.044 | 27.6/0.166 | 1.522 (1.056, 2.193) | 0.024 | 57.7/0.005 |
| Sample size|     |             |        |              |             |         |                      |      |              |       |             |       |
| < 1000     | 34  | 7,651/6,381  | 1.298 (1.164, 1.448) | < 0.001 | 66.6/0.000 | 1.317 (1.163, 1.492) | < 0.001 | 61.1/0.000 | 1.512 (1.264, 1.809) | < 25.5/0.096 | 1.239 (1.102, 1.393) | < 51.4/0.000 | 1.620 (1.324, 1.981) | < 35.9/0.024 |
| ≥ 1000     | 6   | 3,095/8,421  | 0.920 (0.834, 1.015) | 0.097 | 23.9/0.025 | 0.945 (0.841, 1.062) | 0.344 | 255/0.243 | 0.690 (0.539, 0.885) | 0.003 | 0.0/0.758 | 0.981 (0.879, 1.094) | 0.728 | 11.3/0.343 | 0.669 (0.519, 0.862) | 0.002 | 0.0/0661 |
| Quality score|    |             |        |              |             |         |                      |      |              |       |             |       |
| < 10       | 26  | 5,467/4,951  | 1.366 (1.176, 1.586) | < 0.001 | 77.0/0.000 | 1.404 (1.183, 1.667) | < 0.001 | 72.8/0.000 | 1.529 (1.202, 1.944) | 0.001 | 46.0/0.000 | 1.341 (1.121, 1.539) | 0.001 | 64.4/0.000 | 1.692 (1.274, 2.248) | < 58.4/0.000 |
| ≥ 10       | 14  | 5,279/9,851  | 1.038 (0.944, 1.139) | 0.055 | 37.3/0.079 | 1.038 (0.955, 1.128) | 0.376 | 0.0/0.575 | 0.978 (0.719, 1.331) | 0.088 | 49.9/0.017 | 1.043 (0.956, 1.137) | 0.343 | 0.00/818 | 0.985 (0.725, 1.340) | 0.925 | 47.5/0.025 |

n study numbers: Bold values represent statistically significant findings.
Fig. 3 Forest plots of the association between rs2241766 polymorphism and CVD risk. (a) allelic model; (b) dominant model; (c) recessive model; (d) heterozygote model; (e) homozygote model.
Table 4  Overall and subgroup meta-analysis of the association between ADIPOQ rs1501299, +276 G>T polymorphism and CVD

| Categories | Sample size | TVS G OR (95%CI) | TVS G P | TT + TG VS GG OR (95%CI) | TT + TG VS GG P | TG VS GG OR (95%CI) | TG VS GG P |
|------------|-------------|------------------|---------|--------------------------|-----------------|---------------------|-----------|
| Overall    | 44          | 0.956 (0.893, 1.023) | 0.189 | 0.967 (0.890, 1.051) | 0.043 | 0.999 (0.917, 1.051) | 0.098 |
| Population |             |                  |         |                          |                 |                     |           |
| European   | 18          | 0.957 (0.901, 1.016) | 0.146 | 0.967 (0.896, 1.043) | 0.380 | 0.851 (0.717, 1.011) | 0.066 | 35.2/0.070 | 0.988 (0.909, 1.073) | 0.773 | 21.3/0.201 | 0.854 (0.722, 1.012) | 0.068 | 30.2/0.110 |
| East Asian | 20          | 0.966 (0.849, 1.098) | 0.594 | 0.977 (0.834, 1.145) | 0.776 | 0.945 (0.778, 1.149) | 0.572 | 52.2/0.004 | 0.988 (0.858, 1.138) | 0.867 | 64.0/0.000 | 0.940 (0.726, 1.217) | 0.638 | 69.0/0.000 |
| West Asian | 4           | 0.973 (0.643, 1.473) | 0.897 | 0.960 (0.564, 1.635) | 0.880 | 0.999 (0.578, 1.727) | 0.997 | 41.9/0.160 | 0.952 (0.587, 1.546) | 0.843 | 70.2/0.018 | 0.986 (0.477, 2.040) | 0.970 | 62.1/0.048 |
| African    | 2           | 0.848 (0.629, 1.143) | 0.278 | 0.856 (0.583, 1.257) | 0.428 | 0.724 (0.415, 1.266) | 0.258 | 78.0/0.298 | 0.927 (0.614, 1.400) | 0.719 | 0.007/0.725 | 0.700 (0.374, 1.312) | 0.266 | 14.7/0.279 |
| Genotyping |             |                  |         |                          |                 |                     |           |
| PCR-RFLP   | 14          | 0.970 (0.833, 1.128) | 0.688 | 0.997 (0.825, 1.206) | 0.978 | 0.881 (0.684, 1.136) | 0.329 | 55.3/0.006 | 1.051 (0.858, 1.202) | 0.861 | 58.4/0.003 | 0.901 (0.648, 1.253) | 0.535 | 68.4/0.000 |
| TaqMan     | 13          | 0.977 (0.869, 1.099) | 0.701 | 0.987 (0.854, 1.140) | 0.859 | 0.970 (0.791, 1.189) | 0.771 | 301/0.144 | 1.001 (0.874, 1.148) | 0.994 | 47.8/0.028 | 0.956 (0.749, 1.221) | 0.718 | 46.8/0.032 |
| Others     | 17          | 0.930 (0.841, 1.029) | 0.159 | 0.935 (0.827, 1.058) | 0.287 | 0.866 (0.715, 1.048) | 0.140 | 409/0.041 | 0.959 (0.852, 1.079) | 0.484 | 463/0.019 | 0.841 (0.678, 1.044) | 0.117 | 496/0.011 |
| Sample size|             |                  |         |                          |                 |                     |           |
| < 1000     | 36          | 0.945 (0.868, 1.029) | 0.191 | 0.959 (0.884, 1.065) | 0.438 | 0.876 (0.796, 1.016) | 0.079 | 41.8/0.005 | 0.985 (0.896, 1.085) | 0.758 | 47.8/0.001 | 0.965 (0.722, 1.036) | 0.116 | 56.0/0.000 |
| ≥ 1000     | 8           | 0.984 (0.877, 1.014) | 0.784 | 0.985 (0.855, 1.134) | 0.831 | 0.968 (0.784, 1.195) | 0.762 | 44.8/0.080 | 0.987 (0.863, 1.129) | 0.853 | 59.7/0.015 | 0.955 (0.748, 1.219) | 0.711 | 56.3/0.025 |
| Quality score|           |                  |         |                          |                 |                     |           |
| < 10       | 24          | 0.954 (0.848, 1.074) | 0.438 | 0.976 (0.842, 1.132) | 0.752 | 0.879 (0.725, 1.065) | 0.189 | 41.7/0.018 | 1.002 (0.876, 1.145) | 0.981 | 52.8/0.001 | 0.876 (0.683, 1.122) | 0.294 | 59.3/0.000 |
| ≥ 10       | 20          | 0.959 (0.886, 1.038) | 0.298 | 0.963 (0.875, 1.060) | 0.442 | 0.915 (0.782, 1.072) | 0.273 | 44.7/0.017 | 0.976 (0.890, 1.070) | 0.599 | 46.3/0.013 | 0.902 (0.756, 1.075) | 0.250 | 52.2/0.004 |
polymorphism and CVD risk may be due to differences in conflicting results of associations between the ADIPOQ adiponectin protect against this disease [6]. In our meta-analysis, we found that low levels of adiponectin (hypoadipoecytinemia) correlate with the risk of CVD, and high levels of adiponectin protect against this disease [50]. The meta-analysis by Zhao et al. revealed that rs1501299 variant against CVD in general study subjects is significant, which prove our findings will help settle some of the controversies surrounding the ADIPOQ-CVD association research.

**Table 5** Publication bias assessment of this meta-analysis

| SNPs  | Genetic model       | Egger's test | Begg's test |
|-------|---------------------|--------------|-------------|
|       | t-value | P    | z-value | P    |
| rs266729 | Allelic model | 0.60 | 0.52 | 0.47 | 0.639 |
|         | Dominant model     | 0.77 | 0.45 | 0.62 | 0.536 |
|         | Recessive model    | −0.67 | 0.50 | 0.92 | 0.358 |
|         | Heterozygote model | 0.79 | 0.43 | 0.81 | 0.420 |
|         | Homozygote model   | −0.45 | 0.65 | 0.73 | 0.464 |
| rs2241766 | Allelic model | 3.52 | 0.00 | 2.16 | 0.03 |
|          | Dominant model     | 3.63 | 0.00 | 2.99 | 0.00 |
|          | Recessive model    | 0.72 | 0.47 | 0.40 | 0.68 |
|          | Heterozygote model | 3.17 | 0.00 | 2.97 | 0.00 |
|          | Homozygote model   | 0.88 | 0.38 | 0.33 | 0.74 |
| rs1501299 | Allelic model | −0.80 | 0.42 | 0.96 | 0.33 |
|          | Dominant model     | 0.09 | 0.93 | 0.13 | 0.89 |
|          | Recessive model    | −2.24 | 0.03 | 2.11 | 0.03 |
|          | Heterozygote model | 0.60 | 0.54 | 0.11 | 0.91 |
|          | Homozygote model   | −1.45 | 0.15 | 1.49 | 0.13 |

Zhou et al. performed a meta-analysis of the association between rs2241766 and CVD risk in allelic model, and they found that rs2241766 is associated with the increased risk of CVD [63]. In our meta-analysis, we found that rs2241766 is associated with the increased risk of CVD in overall population and East Asian in all the five genetic models, and in West Asian in allelic, recessive, and homozygote models. Our findings is in agreement with the results of Zhou et al., but is in disagreement with the results of Yang et al., Zhang et al., and Zhou et al.

With regard to the association between 1,501,299 and CVD, the results are also conflicting [50, 60, 61]. Zhou et al. revealed no significant association between rs1501299 polymorphism with CAD susceptibility [61]. Qi et al. reported the extremely large decrease in CVD risk associated with rs1501299 polymorphism in diabetic patients [24]. Zhang et al. reported only the weak protective effect of the rs1501299 variant against CAD in general study subjects [50]. The meta-analysis by Zhao et al. revealed that rs1501299 polymorphism may play a protective role for CAD among patients with T2DM [22]. In comparison, our results revealed no significant association.

Different genetic admixture and environmental factors among human populations, which tend to explain ethnic background, strongly modulate the effects of ADIPOQ polymorphisms on adiponectin levels [65, 66]. Studies have reported that low levels of adiponectin (hypoadipocytinemia) correlate with the risk of CVD, and high levels of adiponectin protect against this disease [6–11]. These conflicting results of associations between the ADIPOQ polymorphism and CVD risk may be due to differences in publication bias, sample size, or insufficient statistical power. In addition, evidences have showed that studies which deviate from HWE in controls may reflect the presence of genotyping errors, population stratification, and selection bias in the controls (or without representation of studied sample). Thus, including those studies may decrease the quality of a meta-analysis or generate inconsistent results [67].

Heterogeneity across all the studies of the associations should be noted because it may potentially affect the strengths of the present meta-analysis. We, thus, used random effect model. Our results showed that sample size and quality score are the factors of heterogeneity across all studies of association between rs2241766 polymorphisms and CVD, but no factors contribute the heterogeneity across all studies of association between rs266729/rs1501299 polymorphisms and CVD. However, heterogeneity was still high in the subgroup analysis of the two factors. For these reasons, heterogeneity might be explained by other confounding factors, such as gene-gene interaction and gene-environment interaction.

Our meta-analysis has some limitations. Firstly, significant publication bias was found in the analysis of rs2241766 (under allelic, dominant, and homozygote models) and rs1501299 (under recessive model). Secondly, our meta-analysis mainly included Europeans and Asians with only few other races, thus limiting our power to generalize our findings in other races. Finally, our results might be affected by the potential weaknesses of genetic association studies, such as phenotype misclassifications, genotyping error, population stratification, gene-environment or gene-gene interactions, and selective reporting biases [68, 69].

Despite the limitations highlighted above, our meta-analysis also had some strength. Firstly, we searched extensively and investigated more studies and more participants than any other meta-analyses performed on the association between ADIPOQ variant and CVD, which give our study more statistical power to draw valid conclusion on this issue. Secondly, sensitivity analysis showed that the results of our meta-analysis are stable and robust. Thirdly, the evidence of our results might be explained by other confounding factors, such as gene-gene interaction and gene-environment interaction.

Conclusions

Our meta-analysis found significant increased CVD risk is associated with rs266729 and rs2241766, but not associated with rs1501299. Investigating gene–gene and gene–environment interactions is needed to give more insight into the genetic association between ADIPOQ variants and CVD.
Fig. 4 Sensitivity analyses of the association between rs266729 polymorphism and CVD risk. (a) allelic model; (b) dominant model; (c) recessive model; (d) heterozygote model; (e) homozygote model
Fig. 5 Sensitivity analyses of the association between rs2241766 polymorphism and CVD risk. (a) allelic model; (b) dominant model; (c) recessive model; (d) heterozygote model; (e) homozygote model
Fig. 6 Sensitivity analyses of the association between rs1501299 polymorphism and CVD risk. (a) allelic model; (b) dominant model; (c) recessive model; (d) heterozygote model; (e) homozygote model.
Fig. 7 Trial sequential analysis of the association between rs266729 and CVD risk. (a) allelic model; (b) dominant model; (c) recessive model; (d) heterozygote model; (e) homozygote model
Fig. 8 Trial sequential analysis of the association between rs2241766 and CVD risk. (a) allelic model; (b) dominant model; (c) recessive model; (d) heterozygote model; (e) homozygote model
Fig. 9 Trial sequential analysis of the association between rs1501299 and CVD risk. (a) allelic model; (b) dominant model; (c) recessive model; (d) heterozygote model; (e) homozygote model
Additional files

Additional file 1: Summary of three SNPs characteristics. (DOCX 48 kb)
Additional file 2: Meta-regression results of the association between the SNPs and CVD risk. (DOCX 59 kb)
Additional file 3: Funnel plots of three SNPs for publication bias. (DOCX 1250 kb)
Additional file 4: Additional references. (DOCX 22 kb)

Abbreviations
CAD: Coronary artery disease; CHD: Coronary heart disease; CVD: Cardiovascular disease; HWE: Hardy-Weinberg equilibrium; MI: Myocardial infarction; RIS: Required information size; SNPs: Single-nucleotide polymorphisms; T2DM: Type 2 diabetic mellitus; TSA: Trial sequential analysis

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Please contact author for data requests.

Authors’ contributions
Conception and design: JSK, SQ, YC, and YL. Provision of study materials: JSK, SQ, and RL. Collection and assembly of data: JSK, SQ, and RL. Data analysis and interpretation: JSK, SQ, and RL. Manuscript writing: JSK and SQ. Revised the language/article: All authors. Final approval of manuscript: All authors.

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