Associations between inadequate sleep and obesity in the US adult population: analysis of the national health interview survey (1977–2009)

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Abstract

Background: Epidemiologic studies show a curvilinear relationship between inadequate sleep (< 7 or > 8 hours) and obesity (Body Mass Index > 30 kg/m²), which have enormous public health impact.

Methods: Using data from the National Health Interview Survey, an ongoing nationally representative cross-sectional study of non-institutionalized US adults (≥18 years) (1977 through 2009), we examined the hypothesis that inadequate sleep is independently related to overweight/obesity, with adjustment for socio-demographic, health risk, and medical factors. Self-reported data on health risks, physician-diagnosed medical conditions, sleep duration, and body weight and height were used.

Results: Prevalence of overweight and obesity increased from 31.2% to 36.9% and 10.2% to 27.7%, respectively. Whereas prevalence of very short sleep (<5 hours) and short sleep (5–6 hours) has increased from 1.7% to 2.4% and from 19.7% to 26.7%, it decreased from 11.6% to 7.8% for long sleep. According to multivariate-adjusted multinomial regression analyses, odds of overweight and obesity associated with very short sleep and short sleep increased significantly from 1977 to 2009. Odds of overweight and obesity conferred by long sleep did not show consistent and significant increases over the years. Analyses based on aggregated data showed very short sleepers had 30% greater odds of being overweight or were twice as likely to be obese, relative to 7–8 hour sleepers. Likewise, short sleepers had 20% greater odds of being overweight or 57% greater odds of being obese. Long sleepers had 20% greater odds of being obese, but no greater odds of being overweight.

Conclusions: Our findings support the hypothesis that prevalence of very short and short sleep has gradually increased over the last 32 years. Inadequate sleep was associated with overweight and obesity for each available year.

Keywords: Overweight, Obesity, Short sleep, Long sleep

Background

Inadequate sleep and obesity are important public health challenges with enormous personal, social, and health care costs [1–4]. The extant literature presents mixed findings regarding association between average sleep duration and body weight. A likely explanation for failure to demonstrate such associations may be that the average, rather than inadequate sleep, as herein defined [very short sleep (<5 hours), short sleep (5–6 hours) or long sleep (>8 hours)], is considered as the exposure variable in those studies. It may also be that varying factors such as measurement errors in eliciting responses about habitual sleep duration may have hampered the ability to establish direct comparisons across published studies [5].

The preponderance of epidemiologic evidence suggests a curvilinear relationship between individuals’ habitual sleep duration and their weight, such that inadequate sleep (short or long) is associated with increased risk of obesity among adults in the United States [1,6–11]. Results of longitudinal studies also demonstrated that inadequate sleep
strongly predicts increased body weight [9,10]. A prospective study assessing risk of developing obesity at 1-year among 35,247 respondents sleeping <5 hours or 5–6 hours found adjusted ORs of 1.91 (95% CI 1.36, 2.67) and 1.50 (95% CI 1.24, 1.80), respectively [12]. On balance, there are data suggesting that sleep duration may not be predictive of excess weight [5,8,13,14].

Although available epidemiologic data do not permit ascertainment of cause and effect relationships, they represent a systematic approach for exploring the potential link between inadequate sleep and obesity at the population level. Of particular interest is the observation that prevalence of short sleep and obesity have increased in the past 40 years. Data from the Alameda County Health and Ways of Living Study show a 72% increase in the prevalence of short sleep (<7 hours) from 1965 to 1999 [15]. Data from the National Health and Nutrition Examination Surveys show a 52% increase in the prevalence of obesity from 1971 to 1999 [15]. Increased prevalence of inadequate sleep seems to parallel increased prevalence of obesity for the same period.

Previous studies exploring associations between inadequate sleep and obesity are largely restricted to single years of observation, preventing ascertainment of important trends, or constrained by differing sampling methodologies and analytic approaches, precluding direct comparisons with available studies. One important limitation relates to the use of differing questions to elicit habitual sleep duration across available studies that may yield biased answers [5,16,17]. Such limitations hamper definitive conclusions about the relationship between increased prevalence of obesity and inadequate sleep. The National Health Interview Survey, which has monitored the health of the United States population since 1957, provides a unique data set to determine associations between inadequate sleep and obesity over an extensive period of time using robust analytic techniques. The first hypothesis tested in this study is that both inadequate sleep and overweight/obesity have gradually increased across the last 32 years. The second hypothesis is that inadequate sleep (short or long) is independently associated with overweight and obesity after adjustment for socio-demographic, health-risk, and comorbidity.

**Methods**

The National Health Interview Survey is an ongoing, cross-sectional, in-person household interview survey conducted annually by the Centers for Disease Control’s National Center for Health Statistics. The National Health Interview Survey uses a multistage area probability design, sampling non-institutionalized representatives of the civilian population of the United States. Probability samples of the adult population of all 50 states and District of Columbia were obtained. Details on sample design are provided elsewhere [18].

During face-to-face interviews, respondents provided socio-demographic data (age, sex, race/ethnicity, average family income, and education), health risks (smoking status and alcohol intake), and physician diagnosed chronic conditions or diseases (hypertension, diabetes, coronary heart disease, cancer, kidney disease, stroke, and myocardial infarction). Participants estimated habitual sleep duration (using full-hour increments—5 hours, 6 hours, 7 hours, etc.) by responding to the following: “On average, how many hours of sleep do you get in a 24-hour period?” Sleep duration was coded as either very short sleep (<5 hours/night), short sleep (5–6 hours/night), or long sleep (>8 hours/night), referenced to 7–8 hours/night as healthy sleep [19]. These cut-off points showed significant associations with health risks [20,21].

Body mass index (BMI), which was obtained by computing the ratio of self-reported weight (in kilogram) and height (in meter squared), was coded as overweight (BMI = 25–29.9 kg/m²) and obesity (≥30 kg/m²), referenced to normal weight (BMI = 18.5–24.9 kg/m²). Participants ranked their mood within the last 30 days based on the K-6 scaling system [22]; a score ≥13 indicated emotional distress [23].

Statistical analysis

Analyses were based on National Health Interview Survey data obtained from 1977 through 2009. Since the NHIS dataset includes data from different samples using a multistage area probability sampling design, all analyses performed in this study were based on weighted statistics using the final weights provided with the NHIS dataset. These weights represent a product of weights for corresponding units computed in each of the sampling stages. Once these data were harmonized and standardized across available years, additional processing was unnecessary. As the Center for Disease Control recommends, the complex survey analysis techniques we used accounted for the weights, strata and clusters in the National Health Interview Survey design.

Descriptive analysis was used to ascertain prevalence of variables of interest: overweight/obesity and inadequate sleep across all available years of observation. In order to test the first hypothesis, we used multivariate logistic regression models for complex survey, yielding age and sex-adjusted prevalence of overweight and obesity as well as inadequate sleep (very short sleep, short sleep, and long sleep). Univariate associations of several socio-demographic
(race, education, marital status, poverty), health risk (smoking status, alcohol consumption, physical activity, and emotional distress), and medical conditions (prevalent hypertension, diabetes, stroke, heart attack and heart disease) with these factors were also ascertained. To test the second hypothesis, we used multinomial complex survey logistic regression analysis, assessing the associations of inadequate sleep with overweight/obesity, adjusting for effects of covariates [24,25] for each year with adequate data (1977 – 2009). We also used a hierarchical multivariate-adjusted multinomial logistic regression for complex survey designs to obtain eight separate adjusted models. Since data were not available for each year beginning 1977 and ending 2009, a time-lag factor (three-year interval) was created to adjust for potential time effects on observed associations. Thus, the effect of time was determined using three-year increments. This was achieved by first setting 1977 as the reference year and then dividing the length of time from the reference to the years of data collection by three. Hence, the unit change in time was three years. This approach is similar to a unit of change being one standard deviation on the measurement scale. All analyses were performed using Proc Survey logistic–SAS (SAS Institute, Cary, N.C.) [26,27].

Results
As shown in Table 1, descriptive analyses indicated a trend in the prevalence of inadequate sleep duration (very short, short or long sleep) and overweight/obesity from 1977 to 2009. Specifically, the number of respondents (crude prevalence estimates) classified as overweight or obese increased from 31.2% to 36.9% and 10.2% to 27.7%, respectively. Whereas prevalence of very short sleep (<5 hours) and short sleep (5–6 hours) has increased from 1.7% to 2.4% and from 19.7% to 26.7%, it decreased from 11.6% to 7.8% for long sleep. It is noteworthy that the prevalence of short sleep may have reached a plateau in 2006; with estimates suggesting 27.1% of the United States population were short sleepers.

As shown in Table 2, analyses based on aggregated data (1977–2009) showed very short sleepers had 30% greater odds of being overweight or obese compared with those who slept 7–8 hours. Likewise, short sleepers had 20% greater odds of being overweight or 57% greater odds of being obese. Long sleepers had 20% greater odds of being obese, but no greater odds of being overweight. Results also suggested that age and sex-adjusted prevalence estimates of overweight/obesity and inadequate sleep for the entire study period were significantly affected by several factors. For overweight and obesity, these included self-reported race/ethnicity as black, having less than high school education, being separated, being below the poverty level, being a former or current smoker, former or current drinker, having a diagnosis of diabetes, hypertension, kidney disease, or coronary heart disease, and having experienced a myocardial infarction or a stroke. For very short, short, and long sleep, these factors included black race/ethnicity, being separated, divorced, or widowed, income below the poverty level, being a current or former smoker or drinker, having a diagnosis of kidney disease, diabetes, hypertension, coronary heart disease, or cancer, and having experienced a myocardial infarction or a stroke. Of note, emotional distress was associated with obesity as well as very short, short, and long sleep, while physical activity was associated with obesity and only long sleep.

As shown in Table 3, increased odds of overweight were associated with very short sleep, with odds ranging from 26% to 47%, but results were significant only for years 2004 and 2007; odds of overweight associated with short sleep ranged from 5% to 36%, with significant findings for all years except for 2005 and 2006. Odds of overweight conferred by long sleep ranged from 1% to 11%, but were not significant for any of the available years. Likewise, increased odds of obesity were associated with very short sleep, with odds ranging from 6% to twice as likely, and results were significant for all years except for 1983 and 2006; odds associated with short sleep ranged from 38% to 88%, with significant findings for all years. Odds of obesity conferred by long sleep ranged from 1% to 26%, but were not significant for any of the available years.

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the hypothesis that the prevalence of insufficient sleep (very short or short sleep) has increased over the last 32 years with 1.7% of Americans being very short sleepers in 1977 compared with 2.4% in 2009 and 19.7% of Americans being short sleepers in 1977 compared with 26.7% in 2009.

Table 1 Characteristics of adult respondents in the National Health Interview Survey (1977–2009)

| Characteristics                      | 1977  | 1983  | 1985  | 1990  | 2004  | 2005  | 2006  | 2007  | 2008  | 2009  |
|--------------------------------------|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|
| N                                    | 18,386| 9,931 | 28,999| 35,466| 24,092| 24,290| 18,804| 18,067| 16,926| 21,788|
| Demographics and health risks        |       |       |       |       |       |       |       |       |       |       |
| Gender, female (%)                   | 54.7  | 56.3  | 51.8  | 51.3  | 50.3  | 50.2  | 50.0  | 49.9  | 50.2  | 50.4  |
| Age, years (mean (SD))              | 42(15)| 40(15)| 40(17)| 40(17)| 42(17)| 42(17)| 42(18)| 42(18)| 42(18)| 42(19)| 42(19)|
| Race/ethnicity (%)                  |       |       |       |       |       |       |       |       |       |       |
| Whites                               | 88.6  | 85.7  | 85.7  | 84.0  | 83.4  | 83.0  | 81.5  | 81.4  | 81.2  | 81.0  |
| African Americans/Blacks            | 9.9   | 10.4  | 10.9  | 11.4  | 11.6  | 12.4  | 12.2  | 12.3  | 12.3  | 12.6  |
| Other                                | 1.5   | 3.9   | 3.4   | 4.7   | 4.9   | 5.2   | 6.1   | 6.4   | 6.4   | 6.4   |
| Education, ≥ HS (%)                  | 71.3  | 76.7  | 79.3  | 81.9  | 87.8  | 88.0  | 87.7  | 88.3  | 88.4  | 88.7  |
| Marital status (%)                  |       |       |       |       |       |       |       |       |       |       |
| Married                              | 72.1  | 67.2  | 66.9  | 66.6  | 53.4  | 58.9  | 58.2  | 57.2  | 56.4  | 55.4  |
| Widowed, divorced & separated        | 13.9  | 13.8  | 12.6  | 13.5  | 21.8  | 15.9  | 16.0  | 16.1  | 16.2  | 16.6  |
| Single                               | 14.0  | 19.0  | 20.5  | 19.9  | 24.8  | 25.2  | 25.8  | 26.7  | 27.4  | 28.0  |
| Total family income, ≥ $35 K (%)     | –     | 22.2  | 29.8  | 45.5  | 64.1  | 65.4  | 64.9  | 68.8  | 70.4  | 70.0  |
| Poverty status, below (%)            | –     | 8.5   | 11.2  | 9.4   | 10.2  | 10.4  | 11.7  | 11.3  | 11.6  | 12.5  |
| Ever smoked, yes (%)                 | 58.4  | 55.1  | 55.5  | 50.8  | 42.4  | 42.5  | 41.8  | 41.4  | 42.0  | 42.4  |
| Alcohol consumption (%)              |       |       |       |       |       |       |       |       |       |       |
| Never drinkers                       | –     | –     | –     | 15.3  | 22.4  | 22.1  | 22.6  | 21.3  | 19.0  | 18.2  |
| Former drinkers                      | –     | –     | –     | 8.9   | 13.0  | 12.7  | 12.9  | 13.1  | 12.5  | 13.3  |
| Current drinkers                     | –     | –     | –     | 75.9  | 64.6  | 65.2  | 64.5  | 65.6  | 68.5  | 68.5  |
| Normal (%)                           | 58.6  | 59.4  | 56.6  | 52.6  | 39.3  | 38.4  | 37.7  | 36.9  | 36.5  | 35.4  |
| Overweight (%)                       | 31.2  | 29.1  | 31.6  | 33.4  | 36.9  | 37.1  | 36.2  | 36.8  | 36.0  | 36.9  |
| Obesity (%)                          | 10.2  | 11.4  | 11.8  | 14.0  | 23.8  | 24.6  | 26.1  | 26.2  | 27.5  | 27.7  |
| Physical activity (%)                | –     | –     | –     | –     | 10.9  | 11.5  | 11.2  | 11.5  | 11.9  | 12.7  |
| Emotional distress (%)               | –     | –     | –     | –     | 3.3   | 3.0   | 3.2   | 2.8   | 3.4   | 3.5   |
| Medical conditions                   |       |       |       |       |       |       |       |       |       |       |
| Diabetes (%)                         | –     | 3.6   | –     | –     | 6.3   | 6.7   | 7.0   | 6.8   | 7.5   | 8.2   |
| Hypertension (%)                     | –     | 21.6  | 21.9  | 20.7  | 21.0  | 21.1  | 22.8  | 22.7  | 24.6  | 24.1  |
| Cancer (%)                           | –     | 2.9   | –     | –     | 5.0   | 5.5   | 5.2   | 5.5   | 5.9   | 6.1   |
| Coronary heart disease (%)           | –     | 1.7   | –     | –     | 2.7   | 2.9   | 2.8   | 2.9   | 2.5   | 3.0   |
| Myocardial infarction (%)            | –     | –     | –     | –     | 2.1   | 2.2   | 2.2   | 2.1   | 2.4   | 2.2   |
| Stroke (%)                           | –     | 0.9   | 1.2   | 1.3   | 1.5   | 1.3   | 1.5   | 1.6   | 1.8   | 1.6   |
| Kidney disease (%)                   | –     | –     | –     | –     | 1.3   | 1.3   | 1.2   | 1.1   | 1.2   | 1.4   |
| Sleep measures                       |       |       |       |       |       |       |       |       |       |       |
| Hours of sleep (mean (SD))           | 7.4(1.3)| 7.3(1.3)| 7.4(1.5)| 7.3(1.5)| 7.1(1.5)| 7.1(1.5)| 7.1(1.6)| 7.1(1.6)| 7.1(1.6)| 7.1(1.7)|
| Sleep duration categories (%)        |       |       |       |       |       |       |       |       |       |       |
| Very short sleep                     | 1.7   | 2.0   | 1.3   | 1.5   | 2.2   | 2.1   | 2.1   | 2.4   | 2.0   | 2.4   | 2.4   |
| Short sleep                          | 19.7  | 21.0  | 20.8  | 23.0  | 26.8  | 26.9  | 27.1  | 25.4  | 26.6  | 26.7  |
| Normal sleep                         | 66.8  | 67.5  | 67.1  | 66.6  | 63.7  | 63.6  | 63.5  | 65.6  | 63.5  | 63.0  |
| Long sleep                           | 11.6  | 9.5   | 10.8  | 8.9   | 7.2   | 7.3   | 7.0   | 6.9   | 7.4   | 7.8   |

Abbreviation: SD standard deviation.
Of interest was the comparability of our prevalence estimates for short sleepers (< 7 hours) to those derived from the Sleep Heart Health Study, which was 27.5% in 2005 [32]. Prevalence of short sleep in 1977 was lower than that reported in the data from Alameda County Health and Ways of Living Study, which revealed that 15.2% of respondents slept less than 7 hours in 1965 [15]. Thus, our observation suggests nearly a two-fold increase in the prevalence of short sleepers among Americans, which if replicated in other national datasets, would have an important public health impact. Although fewer Americans reported very short sleep (< 5 hours), it is noteworthy that the prevalence of very short sleep also increased during the same time span. Thus, studies classifying individuals into these two categories may reveal critical health-related outcomes that would be important targets for interventions.

Our findings are consistent with a recent review of trends in sleep duration observed from epidemiologic sleep data in the United States [33]. Accordingly, the preponderance of evidence suggests that United States adults have been sleeping less and less, with the modal sleep duration decreasing from 1979 to 1998 by over one hour [30,33].

### Table 2: Age-sex-adjusted associations of demographics, health risks, and medical conditions with overweight/obesity and inadequate sleep (1977 – 2009)

| Characteristics                  | Overweight (OR 95% CI) | Obesity (OR 95% CI) | Very short sleep (OR 95% CI) | Short sleep (OR 95% CI) | Long sleep (OR 95% CI) |
|----------------------------------|------------------------|---------------------|-----------------------------|------------------------|------------------------|
| Demographics                     |                        |                     |                             |                        |                        |
| Race/ethnicity (reference: Whites) |                        |                     |                             |                        |                        |
| Blacks                           | 1.42 (1.36, 1.47)*     | 1.98 (1.90, 2.06)*  | 2.24 (2.04, 2.45)*          | 1.54 (1.49, 1.60)*     | 1.62 (1.54, 1.71)*     |
| Others                           | 0.73 (0.69, 0.77)      | 0.55 (0.51, 0.60)  | 1.05 (0.88, 1.25)           | 1.15 (1.09, 1.22)     | 0.95 (0.87, 1.05)     |
| Education (reference: High School or higher)* | 1.10 (1.06, 1.14)* | 1.24 (1.19, 1.29)* | 1.93 (1.77, 2.11)*          | 1.04 (1.00, 1.08)*    | 2.01 (1.92, 2.10)*    |
| Marital status (reference: married) |                        |                     |                             |                        |                        |
| Widowed                          | 0.92 (0.87, 0.98)*     | 0.98 (0.91, 1.05)  | 2.17 (1.87, 2.52)           | 1.38 (1.30, 1.47)*    | 1.58 (1.45, 1.73)*    |
| Divorced                         | 0.96 (0.92, 1.00)      | 1.12 (1.07, 1.16)  | 2.67 (2.41, 2.95)           | 1.55 (1.50, 1.61)     | 1.16 (1.09, 1.23)     |
| Separated                        | 1.10 (1.02, 1.18)*     | 1.39 (1.28, 1.50)  | 3.53 (3.03, 4.11)           | 1.59 (1.48, 1.70)     | 1.50 (1.35, 1.66)*    |
| Single                           | 0.81 (0.78, 0.84)      | 0.99 (0.95, 1.03)  | 1.47 (1.32, 1.64)           | 1.03 (0.99, 1.06)     | 1.59 (1.51, 1.67)*    |
| Total family income (reference: ≥ $35,000) | 0.89 (0.87, 0.92)* | 0.95 (0.92, 0.98)* | 2.09 (1.92, 2.27)*           | 1.04 (1.02, 1.07)*    | 1.84 (1.75, 1.92)*    |
| Poverty status (reference: above poverty) | 1.05 (1.00, 1.10)* | 1.39 (1.32, 1.45)* | 2.86 (2.59, 3.17)           | 1.19 (1.14, 1.24)     | 2.04 (1.93, 2.16)*    |
| Smoking status (reference: never smokers) | 0.99 (0.99, 0.99) | 0.99 (0.99, 0.99) | 1.06 (1.05, 1.06)           | 1.02 (1.02, 1.03)*    | 1.04 (1.03, 1.04)*    |
| Alcohol consumption (reference: never) |                        |                     |                             |                        |                        |
| Former                           | 1.09 (1.03, 1.15)*     | 1.35 (1.27, 1.43)  | 1.92 (1.67, 2.21)           | 1.53 (1.45, 1.61)     | 1.35 (1.25, 1.47)*    |
| Current                          | 0.95 (0.92, 0.99)      | 0.86 (0.82, 0.90)  | 0.98 (0.88, 1.10)           | 1.28 (1.23, 1.33)     | 0.82 (0.77, 0.87)     |
| Physical activity (reference: yes)* | 1.05 (0.99, 1.10) | 1.32 (1.24, 1.40)* | 1.08 (0.99, 1.17)           | 1.08 (1.03, 1.18)     | 1.47 (1.32, 1.63)*    |
| Emotional distress (reference: yes)* | 1.07 (0.96, 1.18) | 1.88 (1.71, 1.35)  | 5.51 (5.02, 6.05)           | 5.51 (5.02, 6.05)    | 3.07 (2.72, 3.46)*    |
| Self-reported comorbid conditions |                        |                     |                             |                        |                        |
| Cancer                           | 1.05 (0.99, 1.10)      | 1.32 (1.24, 1.40)  | 1.50 (1.28, 1.77)           | 1.10 (0.99, 1.17)     | 1.47 (1.32, 1.63)*    |
| Diabetes                         | 1.84 (1.69, 1.99)*     | 4.69 (4.43, 5.05)  | 2.15 (1.85, 2.50)           | 1.24 (1.16, 1.31)     | 1.93 (1.77, 2.11)*    |
| Hypertension                     | 1.70 (1.65, 1.76)*     | 3.66 (3.52, 3.79)  | 2.14 (1.95, 2.35)           | 1.28 (1.24, 1.32)     | 1.52 (1.45, 1.60)*    |
| Kidney disease                   | 1.02 (0.87, 1.18)*     | 1.52 (1.31, 1.77)  | 3.10 (2.38, 4.01)           | 0.60 (0.52, 0.68)     | 1.77 (1.47, 2.13)*    |
| Myocardial infarction            | 1.19 (1.04, 1.35)*     | 1.96 (1.73, 2.23)  | 2.89 (2.33, 2.59)           | 1.24 (1.11, 1.38)     | 2.54 (2.20, 2.94)*    |
| Coronary heart disease           | 1.27 (1.13, 1.42)*     | 1.92 (1.72, 2.14)  | 2.30 (1.89, 2.80)           | 1.15 (1.04 1.25)     | 2.15 (1.87, 2.44)*    |
| Stroke                           | 1.07 (0.96, 1.20)      | 1.55 (1.38, 1.74)  | 4.19 (3.45, 5.09)           | 1.39 (1.25, 1.55)     | 3.04 (2.67, 3.45)*    |
| Sleep measure (reference: healthy sleep) |                        |                     |                             |                        |                        |
| Very short sleep                 | 1.30 (1.19, 1.43)*     | 2.17 (1.97, 2.40)*| –                           | –                      | –                      |
| Short sleep                      | 1.20 (1.16, 1.23)*     | 1.57 (1.52, 1.63)*| –                           | –                      | –                      |
| Long sleep                       | 0.99 (0.94, 1.04)      | 1.20 (1.14, 1.26)*| –                           | –                      | –                      |

**Abbreviations:** CI confidence interval, OR odds ratio.

*P < 0.05.

*These measures were available from 2004 to 2009.

*Compares those with less high school education to person with a high school diploma or higher.
Using a different criterion for defining short sleep (<6 hours), investigators noted an increase in short sleep from 7.6% in 1975 to 9.3% in 2006 [34]. Evidently, discrepancy in these two analyses relates to differing definitions of short sleep, which should be considered when comparing prevalence estimates of short sleep across epidemiologic studies [29]. Notwithstanding this discrepant finding, trends observed in both studies are similar, thus supporting the notion that the number of American short sleepers has increased in the last 40 years.

It is noteworthy that the prevalence of short sleep may have reached a plateau in 2006 when 27.1% of United States adults were short sleepers compared with 26.7% in 2009. This observation could be important, but has not been fully elucidated. A similar observation was noted in two previous analyses based on other United States samples. While the first one focusing on data obtained between 1985 and 2006 revealed an increase in the prevalence of short sleep [35], the second one, gathered between 2004 and 2007, showed a decrease in prevalence estimates [36]. It is of interest to examine whether trends observed in our analyses of the National Health Interview Survey data from 1977 to 2009 could be replicated in other epidemiologic sleep data. Future epidemiologic sleep studies are necessary to document any sustained reversal in these observed trends.

If indeed the prevalence of short sleep has ceased to increase, factors underlying this effect should be emphasized in public health campaigns promoting healthy sleep. Plausibly, such factors may provide important clues to help clarify mixed findings in the literature regarding increases in the prevalence of short sleepers over time. Review of the literature suggests that studies that have failed to demonstrate a consistent increase in the number of short sleepers in the United States population have utilized data spanning only the last decade, which is precisely the interval within which the dip seems to have occurred [33]. Those relying on a much longer period of observations have tended to show consistent and gradual increases in the prevalence of short sleep [29,33], as we demonstrated in the present analyses.

Contrary to previous studies that focused only on trends in short sleep, we also ascertained trends in long sleep. As expected, the unadjusted prevalence of long sleep appears to have decreased from 11.6% to 7.8%, which is commensurate with relative increases in the prevalence of very short and short sleep. This could simply be a reflection of gradual decline in sleep duration, affecting all strata of the United States population, as has been demonstrated in previous studies. In effect, examination of data obtained from 1959 to 2005 reveals a decrease in modal sleep duration of approximately 1.5 hours [15,29,30,37]. Epidemiologic data, which have consistently shown a U-shaped association of sleep duration with mortality/morbidity, suggest that whereas increased prevalence of short sleep might have negative effects, reductions in long sleep might have many positive effects [38].

The finding of an increased prevalence of obesity in the US population is consistent with several published papers using other well-characterized epidemiologic studies [3,39]. Specifically, crude prevalence estimates of overweight and obesity increased from 31.2% and 10.2% in 1977 to 36.9% and 27.7% in 2009. Of note, whereas the prevalence of very short and short sleep might have slightly declined after 2007 in the present data, the prevalence of overweight and obesity has not decreased at all. That the continued rise in overweight and obesity from 2004–2009 occurred in the absence of an increase in short sleep duration during the same time period reinforces that the obesity epidemic is probably attributable to multiple other factors, as well as perhaps changes in sleep duration.

### Table 3 Multivariate-adjusted associations of very short, short, and long sleep with overweight/obesity

| Year | Very short sleep | Short sleep | Long sleep | Year | Very short sleep | Short sleep | Long sleep |
|------|-----------------|-------------|------------|------|-----------------|-------------|------------|
| 1983 | 1.37 (0.93, 2.02) | 1.36 (1.19, 1.54)* | 1.02 (0.85, 1.23) | 2003 | 1.34 (0.82, 2.20) | 1.38 (1.15, 1.65)* | 1.01 (0.78, 1.31) |
| 1985 | 1.29 (0.96, 1.75) | 1.18 (1.09, 1.28)* | 0.98 (0.86, 1.10) | 2004 | 1.81 (1.29, 2.54)* | 1.54 (1.38, 1.73)* | 1.08 (0.92, 1.27) |
| 1990 | 0.89 (0.69, 1.16) | 1.31 (1.22, 1.41)* | 1.09 (0.97, 1.22) | 2005 | 1.43 (1.07, 1.91)* | 1.50 (1.37, 1.65)* | 1.04 (0.90, 1.21) |
| 2004 | 1.38 (1.02, 1.85)* | 1.35 (1.15, 1.58)* | 1.11 (0.94, 1.32) | 2006 | 1.92 (1.43, 2.57)* | 1.88 (1.59, 2.21)* | 1.26 (1.05, 1.51) |
| 2005 | 1.26 (0.92, 1.73) | 1.05 (0.93, 1.17) | 1.06 (0.87, 1.30) | 2007 | 2.10 (1.54, 2.86)* | 1.48 (1.31, 1.68)* | 1.19 (0.96, 1.47) |
| 2006 | 1.26 (0.92, 1.73) | 1.05 (0.93, 1.17) | 1.05 (0.86, 1.29) | 2008 | 1.06 (0.87, 1.30) | 1.48 (1.31, 1.68)* | 1.19 (0.96, 1.47) |
| 2007 | 1.47 (1.05, 2.05)* | 1.18 (1.06, 1.32)* | 1.06 (0.87, 1.30) | 2009 | 1.71 (1.21, 2.42)* | 1.40 (1.24, 1.58)* | 1.21 (0.98, 1.49) |
| 2008 | 1.27 (0.90, 1.80) | 1.24 (1.11, 1.39)* | 0.87 (0.72, 1.06) | 2009 | 1.53 (1.08, 2.18)* | 1.38 (1.23, 1.55)* | 0.99 (0.81, 1.21) |
| 2009 | 1.33 (0.97, 1.82) | 1.20 (1.09, 1.34)* | 1.01 (0.84, 1.22) | 2009 | 1.64 (1.18, 2.30)* | 1.55 (1.39, 1.73)* | 1.18 (0.97, 1.42) |

Abbreviations: CI confidence interval, OR odds ratio.
*P < 0.05.
Covariate adjusted in the model included: age, sex, race/ethnicity, education, marital status, family income, poverty status, history of smoking, alcohol consumption, physical activity, emotional stress, history of diabetes, history of hypertension, history of stroke, and coronary heart disease/myocardial infarction.
Of note, a more recent analysis of National Health and Nutrition Examination Survey 2005–2010 data obtained from 13,742 participants (ages ≥20 years) found that short sleepers were more likely to be obese and have abdominal obesity [40]. Although cross-sectional epidemiologic data cannot address causality, the purpose of our analyses was to determine whether inadequate sleep and overweight/obesity were independently associated. The first analytic model ascertained age and sex-adjusted prevalence estimates of inadequate sleep and overweight/obesity for the entire study period (1977–2009), while adjusting potential effects of three clusters of confounders (socio-demographics, health risks, and medical diseases). Such adjustments were necessary because preliminary analyses demonstrated that many of these factors have changed over time (Table 1), with potential adverse effects on both inadequate sleep and overweight/obesity [13,15,29]. Initial selection of specific confounders was guided by results of the Alameda County study showing strong associations with inadequate sleep [15]. A recent review paper provides a comprehensive list of the factors potentially affecting associations between inadequate sleep and obesity [29].

Most of the factors that have contributed to increased odds of overweight/obesity also exhibited positive associations with very short, short, or long sleep. These results underlie the inherent complexity in delineating potential effects of these factors on associations between inadequate sleep and overweight/obesity, which have also changed over time. Intervention aimed at manipulating sleep to modify risk of becoming overweight or obese may require a careful consideration of all of these factors [13,29,41].

| Characteristics | Model 1a | Model 2b | Model 3c |
|-----------------|----------|----------|----------|
| Sleep measures  |          |          |          |
| Reference (7–8 hours) | 1.00 | 1.00 | 1.00 |
| Very short sleep | 1.26 (1.14, 1.38)* | 2.05 (1.85, 2.26)* | 1.29 (1.16, 1.43)* | 1.93 (1.73, 2.16)* | 1.30 (1.16, 1.47)* | 1.95 (1.73, 2.20)* |
| Short sleep     | 1.17 (1.14, 1.20)* | 1.51 (1.46, 1.56)* | 1.19 (1.15, 1.23)* | 1.52 (1.46, 1.58)* | 1.20 (1.15, 1.24)* | 1.53 (1.47, 1.60)* |
| Long sleep      | 1.03 (0.98, 1.08) | 1.30 (1.23, 1.37)* | 1.03 (0.97, 1.09)* | 1.23 (1.15, 1.31)* | 1.04 (0.98, 1.11) | 1.24 (1.16, 1.33)* |

| Characteristics | Model 4d | Model 5e | Model 6f |
|-----------------|----------|----------|----------|
| Sleep measures  |          |          |          |
| Reference (7–8 hours) | 1.00 | 1.00 | 1.00 |
| Very short sleep | 1.25 (1.11, 1.41)* | 1.72 (1.53, 1.95)* | 1.34 (1.18, 1.53)* | 1.91 (1.68, 2.18)* | 1.36 (1.19, 1.55)* | 1.99 (1.75, 2.27)* |
| Short sleep     | 1.18 (1.14, 1.23)* | 1.48 (1.42, 1.55)* | 1.17 (1.12, 1.22)* | 1.51 (1.44, 1.58)* | 1.18 (1.13, 1.23)* | 1.51 (1.45, 1.58)* |
| Long sleep      | 1.02 (0.95, 1.09) | 1.16 (1.08, 1.25) | 1.02 (0.95, 1.10) | 1.20 (1.11, 1.29)* | 1.03 (0.95, 1.11) | 1.24 (1.15, 1.34)* |

| Characteristics | Model 7g | Model 8h |
|-----------------|----------|----------|
| Sleep measures  |          |          |
| Reference (7–8 hours) | 1.00 | 1.00 |
| Very short sleep | 1.30 (1.15, 1.46)* | 1.94 (1.72, 2.19)* | 1.23 (1.08, 1.40)* | 1.59 (1.40, 1.81)* |
| Short sleep     | 1.20 (1.15, 1.24)* | 1.53 (1.47, 1.60)* | 1.14 (1.09, 1.19)* | 1.42 (1.36, 1.49)* |
| Long sleep      | 1.04 (0.98, 1.11) | 1.24 (1.15 1.33)* | 0.99 (0.92, 1.07) | 1.12 (1.03, 1.21)|

Abbreviations: CI confidence interval, OR odds ratio; *P < 0.05.

a Model 1: age-sex-adjusted independent associations of inadequate sleep (short and long) with obesity.
b Model 2: Model 1 plus demographic factors.
c Model 3: Model 2 plus smoking and alcohol consumption.
d Model 4: Model 3 plus hypertension.
e Model 5: Model 3 plus diabetes.
f Model 6: Model 3 plus coronary heart disease/myocardial infarction.
g Model 7: Model 3 plus stroke.
h Model 8: Parsimonious Model (covariates adjusted in the model included: age, sex, race/ethnicity, education, total family income, smoking status, history of diabetes, and history of hypertension).
population [7,9-11,28,42-45]. Indeed, compared with healthy adult sleepers (7–8 hours) a recent meta-analysis showed that the pooled odds ratio for obesity conferred by short sleep was 55% [28]. That we observed smaller odds for obesity may be a function of more adequate adjustment for confounders in our analyses. In all, these epidemiologic findings are in accordance with several studies showing adverse effects of short sleep on body weight [46-49].

Consistent with emerging interest in studying negative effects of long sleep [50], our analyses revealed increased odds of obesity conferred by long sleep; of interest, odds were not greater for overweight. The fact that odds ratios observed for each year were not significant may be explained by reduced cell sizes due to a relatively smaller number of long sleepers and adjustment for multiple factors. By contrast with other studies, 38% of long sleepers were likely to be overweight or obese in the Quebec Family study [51]. Evidently, that study captured a greater proportion of individuals as their outcome variable aggregated both overweight and obesity. Notwithstanding such differences, these findings demonstrate that long sleep may also be a modifiable risk factor; but the evidence linking long sleep with obesity has generally been less robust and consistent than the evidence linking short sleep with obesity. We should note that a recent study analyzing 38 nationally representative time-use surveys conducted across 10 industrialized countries suggests that long sleep might be more widespread than short sleep duration [14], supporting the notion that long sleep may be equally as important as short sleep in devising public health policies.

The major strengths of our study include the use of epidemiologic data from a well-characterized and nationally representative survey of US adults providing adequate data from 1977 to 2009. This has allowed the determination of secular trends of inadequate sleep as well as overweight and obesity and assessment of independent associations between these health indices. Notably, the study has some limitations. First, although our multivariate models adjusted effects of important confounders, causality could not be established as the exposure (inadequate sleep) and outcome (overweight/obesity) measures were collected concurrently. We acknowledge the possibility of bi-directionality in our epidemiologic approach [52], but available experimental evidence supports the hypothesis that inadequate sleep is likely a ‘sufficient cause’ of obesity [53]. Obesity in and of itself does not seem to make substantial contribution to inadequate sleep, as evidence suggests no difference in sleep parameters between obese and non-obese individuals without comorbidity [54]. Second, derivation of inadequate sleep and overweight/obesity was performed on the basis of self-reported data, which may on occasion be subject to varying biases (e.g., poor recall, social desirability, and forward telescoping). We are aware that subjective reports of sleep duration are often overestimated, relative to objectively determined sleep durations [55]. However, we are not aware of any representative studies that have utilized objective sleep recording devices (e.g., polysomnography or actigraphy) permitting analyses of sleep trends in the United States population over a 32-year period [29,41]. Likewise, use of advanced techniques (e.g., bioelectrical impedance, skin-fold thickness, densitometry, computerized tomography, or magnetic resonance imaging) to assess obesity is not feasible in the field [29,41]. We note that self-reported sleep and weight measures correspond well with objective data. Correlation coefficients between subjective and actigraphic sleep durations ranged from 0.57 to 0.78 [56-58]; correlation coefficient between subjective and measured weight was 0.96 [59]. Third, important factors such as sleep apnea, insomnia, and excessive daytime sleepiness or fatigue, which have adverse effects on sleep duration, were not adjusted in our models; such data were not available in the National Health Interview Survey data. However, it bears noting that when previous analyses adjusted effects of such factors, associations between inadequate sleep and obesity remained virtually unaffacted [28,29,41].

Future studies should explore the potential influence of various technological advances (e.g., 24-hour cable television programming, internet use, and smart phones) that have occurred over the last two decades on sleep curtailment and their direct or indirect effect on obesity.

Conclusions

While our analyses were performed on a large dataset with adequate statistical control, it is premature to offer a definitive conclusion regarding causal associations between sleep duration and overweight/obesity. Our study supports the growing evidence that inadequate sleep duration may be an independent risk factor of overweight/obesity. It also appears that the prevalence of overweight/obesity has occurred in tandem with changes in inadequate sleep, which has increased over the last 32 years, although a plateau may have been reached. The findings demonstrate that sleep duration is one domain that should be considered in public health campaigns geared towards decreasing cardio-metabolic risk. Although the obesity-sleep relationship is not entirely understood, considerations should be given to the underlying mechanisms including biological and physiological processes. Physiologic evidence suggests short sleep may influence weight gain through effects on appetite, physical activity, and/or thermoregulation [45]. Care must be exercised when interpreting epidemiologic self-reported sleep data, as varying factors such as measurement errors, response bias, and reverse causation may confound generalizability of the findings. Since we used data from one single source, it is likely that our analyses of subjective data, as captured in the NHIS, would be less biased than analyses comparing results from studies with differing designs and sampling strategies [5,16,17].
Competing interests

The authors declare that they have no competing interests.

Authors’ contributions

JLG, WNJ, SD, YS, ZF, PA, and OG developed the conceptual model. JLG, WNJ, SD, YS, OG contributed to the results and the discussion section. JLG, YS, OG drafted the manuscript. All authors read and approved the final manuscript.

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