Research Article

The Association between Five Genetic Variants in MicroRNAs (rs2910164, rs11614913, rs3746444, rs11134527, and rs531564) and Cervical Cancer Risk: A Meta-Analysis

Jia Liu, Peng Dong, Liane Zhou, and Shijun Wang

Department of Obstetrics and Gynecology, Xuanwu Hospital, Capital Medical University, Beijing, China

Correspondence should be addressed to Shijun Wang; wsj2018@yeah.net

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The objective of this study was to conduct a meta-analysis to systematically summarize and investigate the association of miRNA-124 rs531564, miRNA-218 rs11134913, miRNA-146a rs2910164, miRNA-196a2 rs11614913, and miRNA-499 rs3746444 polymorphisms with cervical cancer. A systematic review was performed to identify relevant studies using Embase and PubMed databases. A chi-square-based Q-test combined with the inconsistency index ($I^2$) was used to check the heterogeneity between studies. A total of six case-control studies on rs2910164 and rs11614913, 4 studies on rs3746444 and rs11134527, and three studies on rs531564 were included. No evidence of association was found between miR-146a rs2910164, miR-196a2 rs11614913, miRNA-499 rs3746444, and miR-218 rs11134527 polymorphisms and cervical cancer risk in all the genetic models. The miR-124 rs531564 polymorphism was associated with a statistically increased risk of cervical cancer in a homozygote model (CC vs. GG: OR = 2.87, 95% CI: 1.40-5.91, $P_H = 0.887$), dominant model (GC/CC vs. GG: OR = 1.38, 95% CI: 1.07-1.80, $P_H = 0.409$), and recessive model (CC vs. GC/GG: OR = 2.26, 95% CI: 1.58-3.23, $P_H = 0.979$). However, this finding should be interpreted with caution for limited samples and heterogeneity. Large-scale and well-designed studies are needed to validate our result.

1. Introduction

Cervical cancer (CC) is the fourth most common cancer in females, with 570,000 new cases and 311,000 deaths estimated for 2018 worldwide [1]. Cervical cancer is a multistep process involving the transformation of the normal cervical epithelium to cervical intraepithelial neoplasia that is subsequently transformed to cervical cancer [2]. Due to cervical cytology screening, cervical cancer can be detected at an early stage, greatly reducing the incidence and mortality. However, cervical cancer is still one of the deadliest female-specific cancers due to its tendency to metastasize and recur after treatment. It is well established that the persistence of Human Papillomavirus (HPV) infection is the main cause of cervical cancer and is indeed deemed as a necessary cause for the disease. However, evidence suggests that genetic risk factors also play a crucial role in the pathogenesis of cervical cancer [3–5].

MicroRNAs (miRNAs) are short nonprotein-coding small RNAs of approximately 23 nucleotides that mediate posttranscriptional regulation through base pairing to partially complementary sites of the target mRNA and promote translational repression or messenger RNA degradation [6, 7]. Increasing evidences show that single nucleotide polymorphisms (SNPs) or variation in miRNA sequence could potentially alter various biological processes by influencing target selection of miRNAs and is closely associated with various human diseases [8]. The relationship between genetic mutations in microRNA and cancer susceptibility has been getting increasing attention. Several recent studies have indicated that miRNA-related SNPs can remarkably alter the biogenesis and function of the corresponding miRNAs and may function as oncogenic or antioncogenic molecules in human cancers [9, 10].

Recently, though a number of studies have been conducted to investigate the association of miR-124 G>C (rs531564), miR-218 A>G (rs11134527), miR-146a G>C (rs2910164), miR-196a2 T>C (rs11614913), and miR-499 A>G (rs3746444) polymorphisms with cervical cancer [2,
11–20], the results were inconclusive. Therefore, a meta-analysis was performed to evaluate the association between these five SNPs and the risk of cervical cancer.

2. Methods and Materials

2.1. Publication Search. The PubMed and Embase databases were searched by two independent investigators covering all papers published up to December 27, 2020. The following terms were used: "cervical or CIN", "microRNA or miRNA or microRNAs", and "polymorphism or polymorphisms". The references of the retrieved articles and review articles were also searched manually for further relevant studies. All eligible studies met the following criteria: (1) full-text study, (2) case-control design, (3) analysis of the relationship between miRNA polymorphisms and the risk of cervical cancer, (4) studies focusing on human beings, and (5) sufficient data to calculate risk estimates: the odds ratio (OR) with 95% confidence interval (CI) and a P value. The selection of the study was completed independently by two investigators (Liu and Dong) according to the inclusion criteria.

2.2. Data Extraction. Each publication was sought for the following information: the first author’s surname, the publication year, country of origin, ethnicity, genotyping method, the number of cases and controls, and Hardy-Weinberg equilibrium (HWE). According to the amount of the sample, we defined sample size as the large sample size (≥200) and small sample size (<200).

Two reviewers (Liu and Dong) extracted the eligible research data repeatedly. Disagreement between two reviewers was discussed with another reviewer (Wang) until a consensus is reached.

When Hardy-Weinberg equilibrium (HWE) in the control was not reported, the $\chi^2$ test was used to evaluate the HWE of polymorphism in the control group. If $P < 0.05$, it is considered deviated from HWE.

2.3. Methodological Quality Assessment. The quality of the selected studies was assessed according to the methodological quality assessment scale which was mentioned in prior meta-analysis [22] (see Table S1).

In this scale, five items were evaluated, including (1) the representativeness of the subjects, (2) the ascertainment of cervical cancer, (3) the control of the genotyping methods, (4) HWE in controls, and (5) assessment of association. The quality score ranged from 0 to 9, and a high score indicated good quality of the study. Studies with a score less than 5 were removed from the subsequent analyses.

2.4. Statistics. We conducted our meta-analysis based on the PRISMA checklists and followed the guideline [23]. The strength of association was assessed between miRNA polymorphisms and the risk of cervical cancer by crude OR. Pooled ORs were performed for allelic comparison (miR-146a: G versus C, miR-196a: T versus C, miR-499: A versus G, miR-218: A versus G, and miR-124: G versus C), heterozygote model (miR-146a: GC versus GG, miR-196a: TC versus TT, miR-499: AG versus AA, miR-218: AG versus AA, and miR-124: GC versus GG), homozygote model (miR-146a: CC versus GG, miR-196a: CC versus TT, miR-499: GG versus AA, miR-218: GG versus AA, and miR-124: CC versus TT), dominant model (miR-146a: CC+GC versus GG, miR-196a: TC+CC versus TT, miR-499: AG+GG versus AA, miR-218: AG+GG versus AA, and miR-124: GC+CC versus GG), and recessive model (miR-146a: CC versus CG+GG, miR-196a: TC versus TT, miR-499: GG versus AG+AA, miR-218: GG versus AG+AA, and miR-124: CC versus GC+GG), respectively.

A chi-square-based Q-test combined with the inconsistency index ($I^2$) was used to check the heterogeneity between studies [25]. The random-effects model was used when heterogeneity tests yielded significant results ($P < 0.1$ or $I^2 ≥ 50\%$) [26, 27]. Besides, subgroup analyses were stratified by sample size and HWE. Publication bias was assessed by visual inspection of funnel plots and using Egger’s test. Sensitivity analyses were performed to identify individual studies’ effect on pooled results and test the reliability of results. $P$ value of Egger’s test $<0.05$ was considered representative of statistically significant publication bias. All statistical analyses were carried out using the STATA 12.0 software (Stata Corp., College Station, TX, USA).

3. Results

3.1. Study Characteristics. A total of 247 studies were acquired from Embase and PubMed databases (Embase: 98 and PubMed: 149). The detailed screening process is shown in Figure 1. After reading the titles and the abstracts, 235 articles were excluded, of which 61 were duplicate ones, 149 had no relation to this topic, 22 were reviews or meta-analyses, and 3 were conference abstracts. Finally, 12 eligible case-control studies were included in our meta-analysis (Table 1) [10–21].

In six studies, genotype frequencies of rs531564, rs11134527, rs2910164, rs11614913, and rs3746444 were presented separately, so each of them was considered as a separate study in this meta-analysis. Therefore, all the six included studies containing 2291 cases and 2850 controls for rs2910164, six studies containing 2291 cases and 2850 controls for rs11614913, four studies involving 1554 cases and 1962 controls for rs3746444, four studies containing 3167 cases and 3132 controls for rs11134527, and three studies containing 975 cases and 950 controls for rs531564 were finally analyzed in our meta-analysis. For these five miRNA in this meta-analysis, the subjects in all included studies were of the Asian population.

The TaqMan assay and polymerase chain reaction-restriction fragment length polymorphism (PCR-RFLP) were used for genotyping to determine the SNPs in the included studies. HWE of genotype distribution in the controls was tested in all studies, and two articles were not consistent with HWE. And according to the methodology quality assessment, all the studies were above a score of 6.0 and recruited into the following analyses.

3.2. Association between miRNA-146a rs2910164 Polymorphism and Cervical Cancer Susceptibility. The association strength between miRNA-146a rs2910164 polymorphism
and cervical cancer risk is shown in Table 2. As shown in Table 2, significant association was not identified in any genetic model (C vs. G: OR = 1.08, 95% CI 0.90-1.30, \( P_H = 0.001 \), Figure 2; GC vs. GG: OR = 0.94, 95% CI 0.76-1.15, \( P_H = 0.140 \); CC vs. GG: OR = 0.85, 95% CI 0.57-1.26, \( P_H = 0.001 \); GC/GG vs. CC: OR = 0.90, 95% CI 0.70-1.15, \( P_H = 0.014 \); and GG vs. GC/CC: OR = 0.90, 95% CI 0.68-1.19, \( P_H = 0.002 \)).

Next, subgroup analysis was conducted according to sample size and HWE. In the large sample size, heterogeneity is significantly reduced. A significantly increased risk of cervical cancer susceptibility was identified in homozygote comparison (CC vs. TT: OR = 1.19, 95% CI: 1.00-1.41, \( P_H = 0.369 \)). In the HWE group, heterogeneity is also significantly reduced. But significant association was not identified in any genetic model.

### 3.4. Association between miR-499 rs3746444 Polymorphism and Cervical Cancer Susceptibility.

Significant association was not identified in any genetic model (G vs. A: OR = 0.76, 95% CI: 0.50-1.15, \( P_H = 0.000 \), Figure 4; AG vs. AA: OR = 1.34, 95% CI: 0.82-2.20, \( P_H = 0.001 \); GG versus AA: OR = 1.42, 95% CI: 0.61-3.30, \( P_H = 0.000 \); AG/GG vs. AA: OR = 1.42, 95% CI: 0.83-2.42, \( P_H = 0.000 \); and GG vs. AG/AA: OR = 1.27, 95% CI: 0.68-2.35; \( P_H = 0.002 \)).

Next, subgroup analysis was conducted according to sample size and HWE. In the large sample size, heterogeneity is still high, and in a large sample size, the heterogeneity is high except the homozygote model and recessive model. When we delete the two articles (Bin Zhou and Shruti SrivaStava),

3.3. Association between miR-196a2 rs11614913 Polymorphism and Cervical Cancer Susceptibility. Signiﬁcant association was not identiﬁed in any genetic model (T vs. C: OR = 0.83, 95% CI: 0.67-1.04, \( P_H = 0.000 \), Figure 3; TC vs. TT: OR = 1.19, 95% CI: 0.90-1.58, \( P_H = 0.001 \); CC vs. TT: OR = 1.41, 95% CI: 0.94-2.10, \( P_H = 0.000 \); TC/CC vs. TT: OR = 1.28, 95% CI: 0.93-1.75, \( P_H = 0.000 \); and CC vs. TC/TT: OR = 1.23, 95% CI: 0.97-1.29, \( P_H = 0.021 \) when all eligible studies were pooled.

Next, subgroup analysis was conducted according to sample size and HWE. In large sample size, heterogeneity is signiﬁcantly reduced. A signiﬁcantly increased risk of cervical cancer susceptibility was identiﬁed in homozgyote comparison (CC vs. TT: OR = 1.19, 95% CI: 1.00-1.41, \( P_H = 0.369 \)). In the HWE group, heterogeneity is also signiﬁcantly reduced. But signiﬁcant association was not identiﬁed in any genetic model.
heterogeneity is significantly reduced in all models (Table 3). However, significant association was also not identified in any genetic model.

3.5. Association between miR-218 rs11134527 Polymorphism and Cervical Cancer Susceptibility. Significant association was not identified in any genetic model (G vs. A; OR = 1.11, 95% CI: 0.98-1.25). While there is a significantly increased risk of cervical cancer susceptibility in the homozygote model (GG vs. AA: OR = 1.31, 95% CI: 1.14-1.50), the sample size of the 4 included articles is all large size and subgroup analysis was conducted according to HWE. In the HWE group, heterogeneity is also significantly reduced. But significant association was not identified in any genetic model. When we delete two articles (Li Chuanyin and Guange Chen), heterogeneity is significantly reduced in all models (Table 3). We observed a significantly increased risk of cervical cancer susceptibility in allelic comparison (G versus A: OR = 1.11, 95% CI: 1.02-1.21, P = 0.395), while there is a significantly decreased risk of cervical cancer susceptibility in the homozygote model (GG vs. AA: OR = 0.79, 95% CI: 0.66-0.94, P = 0.594) and recessive model (GG vs. AA: OR = 0.80, 95% CI: 0.68-0.94, P = 0.982). Significant association was not identified in heterozygote comparison (AG vs. AA: OR = 0.97, 95% CI: 0.85-1.10, P = 0.214) and dominant model (GG/AG vs. AA: OR = 0.92, 95% CI: 0.81-1.04, P = 0.235).

3.6. Association between miR-124 rs531564 Polymorphism and Cervical Cancer Susceptibility. We observed a significantly increased risk of cervical cancer susceptibility in the homozygote model (CC vs. GG: OR = 2.87, 95% CI: 1.40-
| miRNA     | N    | OR (95% CI) | PH | OR (95% CI) | PH | OR (95% CI) | PH | OR (95% CI) | PH | OR (95% CI) | PH |
|-----------|------|-------------|----|-------------|----|-------------|----|-------------|----|-------------|----|
| miRNA-146a C/G | 6.00 | 1.08 (0.90-1.30) | 0.001 | 0.94 (0.76-1.15) | 0.140 | 0.85 (0.57-1.26) | 0.001 | 0.90 (0.70-1.15) | 0.014 | 0.90 (0.68-1.19) | 0.002 |
| Sample size Large | 4.00 | 1.13 (0.91-1.40) | 0.001 | 0.88 (0.67-1.16) | 0.072 | 0.77 (0.49-1.20) | 0.001 | 0.83 (0.60-1.15) | 0.012 | 0.86 (0.64-1.16) | 0.003 |
| HWE P > 0.05 | 5.00 | 1.06 (0.86-1.31) | 0.000 | 0.93 (0.73-1.19) | 0.083 | 0.88 (0.56-1.39) | 0.000 | 0.91 (0.67-1.22) | 0.007 | 0.93 (0.69-1.27) | 0.001 |
| P < 0.05 | 1.00 | 1.24 (0.88-1.74) | — | 0.91 (0.55-1.52) | — | 0.68 (0.36-1.31) | — | 0.83 (0.53-1.31) | — | 0.71 (0.38-1.32) | — |
| miRNA-196a2 C/T | 6.00 | 0.83 (0.67-1.04) | 0.000 | 1.19 (0.90-1.58) | 0.001 | 1.41 (0.94-2.10) | 0.000 | 1.28 (0.93-1.75) | 0.000 | 1.23 (0.97-1.57) | 0.021 |
| Sample size Large | 4.00 | 0.92 (0.84-1.02) | 0.284 | 1.06 (0.86-1.31) | 0.106 | 1.19 (1.00-1.41) | 0.369 | 1.10 (0.90-1.33) | 0.112 | 1.12 (0.97-1.29) | 0.676 |
| Small | 2.00 | 0.62 (0.22-1.78) | 0 | 1.91 (0.52-7.08) | 0.001 | 2.24 (1.32-15.50) | 0.000 | 2.12 (0.44-10.21) | 0.000 | 1.49 (0.49-4.52) | 0.006 |
| HWE P > 0.05 | 5.00 | 0.93 (0.85-1.02) | 0.336 | 1.06 (0.89-1.26) | 0.182 | 1.16 (0.99-1.37) | 0.394 | 1.08 (0.92-1.28) | 0.173 | 1.10 (0.96-1.27) | 0.681 |
| P < 0.05 | 1.00 | 0.36 (0.26-0.51) | — | 3.81 (1.97-7.37) | — | 5.99 (3.09-11.59) | — | 4.80 (2.62-8.76) | — | 2.57 (1.59-4.15) | — |
| miRNA-499 G/A | 4.00 | 0.76 (0.50-1.15) | 0 | 1.34 (0.82-2.20) | 0.001 | 1.42 (0.61-3.30) | 0 | 1.42 (0.83-2.42) | 0 | 1.27 (0.68-2.35) | 0.002 |
| Sample size Large | 2.00 | 0.86 (0.58-1.27) | 0.025 | 1.34 (0.65-2.75) | 0.001 | 1.03 (0.66-1.60) | 0.705 | 1.28 (0.70-2.35) | 0.003 | 0.99 (0.64-1.53) | 0.413 |
| Small | 2.00 | 0.68 (0.27-1.72) | 0 | 1.34 (0.52-3.48) | 0.035 | 2.01 (0.35-11.56) | 0 | 1.58 (0.44-5.63) | 0.002 | 1.67 (0.55-5.05) | 0.001 |
| HWE P > 0.05 | 2.00 | 0.67 (0.28-1.58) | 0 | 1.36 (0.61-3.00) | 0.007 | 2.28 (0.52-9.98) | 0 | 1.64 (0.54-5.01) | 0 | 1.80 (0.69-4.74) | 0.006 |
| P < 0.05 | 2.00 | 0.86 (0.55-1.36) | 0.048 | 1.33 (0.56-3.16) | 0.029 | 0.84 (0.50-1.43) | 0.886 | 1.26 (0.59-2.70) | 0.035 | 0.89 (0.60-1.34) | 0.583 |
| miRNA-218 G/A | 4.00 | 0.98 (0.83-1.15) | 0.003 | 1.07 (0.89-1.29) | 0.053 | 1.01 (0.73-1.40) | 0.006 | 1.06 (0.86-1.31) | 0.010 | 0.97 (0.76-1.23) | 0.035 |
| Overall | 4.00 | 0.98 (0.83-1.15) | 0.003 | 1.07 (0.89-1.29) | 0.053 | 1.01 (0.73-1.40) | 0.006 | 1.06 (0.86-1.31) | 0.010 | 0.97 (0.76-1.23) | 0.035 |
| Sample size Large | 4.00 | 0.98 (0.83-1.15) | 0.003 | 1.07 (0.89-1.29) | 0.053 | 1.01 (0.73-1.40) | 0.006 | 1.06 (0.86-1.31) | 0.010 | 0.97 (0.76-1.23) | 0.035 |
| Small | 0.00 | — | — | — | — | — | — | — | — | — | — |
| HWE P > 0.05 | 3.00 | 1.03 (0.88-1.21) | 0.019 | 0.99 (0.88-1.13) | 0.318 | 0.92 (0.66-1.28) | 0.021 | 0.97 (0.82-1.16) | 0.091 | 0.92 (0.69-1.24) | 0.031 |
| P < 0.05 | 1.00 | 0.79 (0.63-0.98) | — | 1.53 (1.09-2.15) | — | 1.43 (0.94-2.16) | — | 1.49 (1.09-2.05) | — | 1.15 (0.79-1.66) | — |
| miRNA-124 C/G | 3.00 | 0.58 (0.40-0.83) | 0.097 | 1.27 (0.97-1.66) | 0.899 | 2.87 (1.40-5.91) | 0.887 | 1.38 (1.07-1.80) | 0.409 | 2.26 (1.58-3.23) | 0.979 |
Table 2: Continued.

| Sample size | N  | OR   | PH | OR    | PH | OR    | PH | OR    | PH | OR    | PH | OR    | PH |
|-------------|----|------|----|-------|----|-------|----|-------|----|-------|----|-------|----|
| Large       | 2  | 0.46 (0.32-0.66) | 0.858 | 1.70 (0.47-6.17) | 0.967 | 3.63 (1.04-12.63) | 0.996 | 3.13 (0.90-10.85) | 0.995 | 2.24 (1.51-3.31) | 0.860 |
| Small       | 1  | 0.75 (0.59-0.95) | —    | 1.25 (0.95-1.64) | —    | 2.48 (1.02-6.04) | —    | 1.32 (1.01-1.72) | —    | 2.36 (0.97-5.74) | —    |
| HWE         |    |       |      |       |      |       |      |       |      |       |      |       |      |
| P > 0.05    | 3  | 0.58 (0.40-0.83) | 0.097 | 1.27 (0.97-1.66) | 0.899 | 2.87 (1.40-5.91) | 0.887 | 1.38 (1.07-1.80) | 0.035 | 2.26 (1.58-3.23) | 0.979 |
| P < 0.05    | 0  | —     | —    | —     | —    | —     | —    | —     | —    | —     | —    | —     | —    |
5.91, \( P_H = 0.887 \), Figure 6), dominant model (GC/CC vs. GG: \( OR = 1.38, \) 95% CI: 1.07-1.80, \( P_H = 0.409 \)), and recessive model (CC vs. GC/GG: \( OR = 2.26, \) 95% CI: 1.58-3.23, \( P_H = 0.979 \)), while there is a significantly decreased risk of cervical cancer susceptibility in allelic comparison (C vs. G: \( OR = 0.58, \) 95% CI: 0.40-0.83, \( P_H = 0.097 \)). Significant association was not identified in heterozygote comparison (GC vs. GG: \( OR = 1.27, \) 95% CI: 0.97-1.66, \( P_H = 0.899 \)).

Next, subgroup analysis was conducted according to sample size. In a large sample size, a significantly increased risk of cervical cancer susceptibility could be identified in the homozygote model (CC vs. GG: \( OR = 3.63, \) 95% CI: 1.04-12.63, \( P_H = 0.998 \)) and recessive model (CC vs. GC/GG: \( OR = 2.24, \) 95% CI: 1.51-3.31, \( P_H = 0.86 \)), while there is a significantly decreased risk of cervical cancer susceptibility in allelic comparison (C vs. G: \( OR = 0.46, \) 95% CI: 0.32-0.66, \( P_H = 0.858 \)). Significant association was not identified in heterozygote comparison (GC vs. GG: \( OR = 1.27, \) 95% CI: 0.97-1.66, \( P_H = 0.899 \)) and dominant model (GC/CC vs. GG: \( OR = 3.13, \) 95% CI: 0.90-10.85, \( P_H = 0.995 \)).

3.7. Sources of Heterogeneity. Since significant heterogeneity was found for miR-146 (rs2910164), miR-196a2 (rs11614913), miR-499 (rs3746444), and miR-218 (rs11134527), we conducted a subgroup analysis by sample size, and the results showed that the sample size is the source of heterogeneity of miR-196a2 (rs11614913) in each genetic model. However, for other SNPs, heterogeneity was still high in large sample size. However, when we delete some articles, we found that the heterogeneity is significantly reduced.

3.8. Sensitivity Analysis. The influence of each study on the pooled ORs in this meta-analysis was examined by excluding one study at a time using STATA 12.0 software. The results showed that no significant alteration in the pooled ORs was found in any of the genetic models for the five SNPs.

3.9. Publication Bias. Due to the limited number of studies included (<10), no funnel plot or Egger’s test was performed.
### Table 3: Summary of pooled ORs after deleting some articles.

| miRNA    | N       | OR   | PH   | OR   | PH   | OR   | PH   | OR   | PH   |
|----------|---------|------|------|------|------|------|------|------|------|
| miRNA-146a | Overall | 6    | 1.08 (0.90-1.30) | 0.001 | 0.94 (0.76-1.15) | 0.140 | 0.85 (0.57-1.26) | 0.001 | 0.90 (0.70-1.15) | 0.014 | 0.90 (0.68-1.19) | 0.002 |
| miRNA-499 | After deleting three articles | 3    | 1.36 (1.19-1.56) | 0.800 | 0.76 (0.60-0.97) | 0.379 | 0.53 (0.40-0.70) | 0.683 | 0.68 (0.54-0.85) | 0.406 | 0.66 (0.53-0.82) | 0.638 |
| miRNA-218 | Overall | 4    | 0.76 (0.50-1.15) | 0.000 | 1.34 (0.82-2.20) | 0.001 | 1.42 (0.61-3.30) | 0.000 | 1.42 (0.83-2.42) | 0.000 | 1.27 (0.68-2.35) | 0.002 |
| miRNA-218 | After deleting two articles | 2    | 1.04 (0.90-1.21) | 0.730 | 0.93 (0.77-1.12) | 0.674 | 0.97 (0.65-1.46) | 0.514 | 0.94 (0.79-1.13) | 0.633 | 1.01 (0.72-1.42) | 0.673 |
| miRNA-218 | Overall | 4    | 0.98 (0.83-1.15) | 0.003 | 1.07 (0.89-1.29) | 0.053 | 1.01 (0.73-1.40) | 0.006 | 1.06 (0.86-1.31) | 0.010 | 0.97 (0.76-1.23) | 0.035 |
| miRNA-218 | After deleting two articles | 2    | 1.11 (1.02-1.21) | 0.395 | 0.97 (0.85-1.10) | 0.214 | 0.79 (0.66-0.94) | 0.594 | 0.92 (0.81-1.04) | 0.235 | 0.80 (0.68-0.94) | 0.982 |
4. Discussion

In this meta-analysis, the relationship between these five SNPs in microRNAs (rs2910164, rs11614913, rs3746444, rs11134527, and rs531564) and the risk of cervical cancer was evaluated from 11 published studies. We demonstrated that no evidence was found for the association between miRNA-146a rs2910164, miR-196a2 rs11614913, miRNA-499 rs3746444, and miR-218 rs11134527 and cervical cancer risk in any genetic models. The miR-124 rs531564 polymorphism was associated with a statistically increased risk of cervical cancer in the homozygote model, dominant model, and recessive model. However, due to limited samples and heterogeneity, this finding should be interpreted with caution.

Recent studies have demonstrated that miRNAs are abnormally expressed in cervical cancer and play a role in tumor generation [28]. miR-146a (rs2910164) involves a G>C nucleotide substitution, which changes the G:U pair to C:U [29]. Different locations may cause the function of miR-146a to be different, and the binding affinity to the 3′-UTR region of the target gene may also be different [30, 31]. Furthermore, it is reported that miR-146a can restrict the NF-κB signaling pathway mediated by IRAK1 and TRAF6, which is an important signaling factor involved in the occurrence of cancers [32, 33]. Previous studies have shown that miR-124 is closely related to the occurrence and development of cervical cancer [34, 35]. Evidence has been shown that miR-124 rs531564 polymorphism affects miRNA processing, and the G allele affects the circular secondary structure of pre-miR-124 and changes the expression of other miRNAs [36].

In 2013, Yamamoto et al. found that the expression of miR-218 in cervical cancer tissues was significantly lower than that of adjacent noncancerous tissues. The restoration of miR-218 significantly inhibited the proliferation, migration, and invasion of cancer cells [37]. Laminin-5 β3 (LAMB3) has been verified as a transcriptional target of miR-218 [38], and the expression of LAMB3 is increased in the presence of the HPV-16 E6 oncoprotein, and this effect is mediated by miR-218 [19]. The miR-196a2 is located on the human chromosome 12q13.13. In 2016, Torruella-Loran et al. reported that rs11614913 in miR-196a2 has the function of regulating the expression of several genes involved in cancer [39]. Results of a functional study showed...
that the mutation reduced the efficiency of miRNA precursor processing into its mature form and reduced the ability to regulate target genes [40]. miR-499 is one of the microRNAs that play an important role in posttranscriptional regulation by regulating multiple genes and signal transduction pathways. The miR-499 (rs3746444) polymorphism leads to mismatches from A:U to G:U in the stem structure of the miR-499 precursor. The presence of this mismatch will affect Sox6 and Rod1 genes, which are important for the occurrence of cancers [41, 42].

Significant heterogeneity was found in this meta-analysis for association of miR-218 rs11134527, miR-146a rs2910164, miR-196a2 rs11614913, miR-499 rs3746444, and cervical cancer. Therefore, we conducted a subgroup analysis based on sample size and HWE. When we restricted the sample size to the large sample size group and the HWE group, heterogeneity is significantly reduced for miR-196a2 rs11614913. However, the small sample size group of miR-196a2 includes only two studies, and the large sample size group includes three studies. Therefore, further research is needed to confirm these results. Significant heterogeneity was also found for the association of miRNA-146a rs2910164, miR-499 rs3746444, miR-218 rs11134527, and cervical cancer in the large sample size group and the HWE group. However, when we delete some researches in each analysis, heterogeneity is significantly reduced, suggesting that these researches contributed to the source of heterogeneity. However, owing to the limited articles included, further evaluation of the results is still needed.

In the sensitivity analysis, no significant changes were found after omitting one study at a time, indicating that our meta-analysis results are relatively stable and credible. Our meta-analysis has several advantages. First, this is the first meta-analysis of the association between five miRNA polymorphisms and the susceptibility to cervical cancer. In addition, according to the methodological quality assessment, all included studies are of high quality. Moreover, there are no restrictions in the literature search, so the selection bias is well controlled. Our meta-analysis has some limitations. Firstly, the number of included studies is too few. Secondly, heterogeneity was detected in miR-146a rs2910164, miR-196a2 rs11614913, miR-499 rs3746444, and miR-218 rs11134527. After deleting some studies in each analysis, heterogeneity is significantly reduced. The pooled ORs became significant without evidence of heterogeneity. Thirdly, the included studies were all Asian ethnicity; therefore, it is uncertain whether these results can be generalized to other populations. Fourth, publication bias might also

| Study ID                  | OR (95% CI)   | % weight |
|--------------------------|---------------|----------|
| Large size               |               |          |
| Bin Zhou (2011)          | 0.69 (0.51, 0.94) | 24.54   |
| Shizhi Wang (2019)       | 1.03 (0.87, 1.21) | 27.15   |
| Subtotal ($I^2 = 80.2\%, p = 0.025$) | 0.86 (0.58, 1.27) | 51.69   |

| Small size               |               |          |
| Shruti Srivastava (2017) | 0.43 (0.31, 0.58) | 24.59   |
| Nisha Thakur (2019)      | 1.10 (0.78, 1.55) | 23.72   |
| Subtotal ($I^2 = 93.8\%, p = 0.000$) | 0.68 (0.27, 1.72) | 48.31   |
| Overall ($I^2 = 89.4\%, p = 0.000$) | 0.76 (0.50, 1.15) | 100.00  |

NOTE: Weights are from random effects analysis

(a)

| Study ID                  | OR (95% CI)   | % weight |
|--------------------------|---------------|----------|
| $P < 0.05$               |               |          |
| Bin Zhou (2011)          | 0.69 (0.51, 0.94) | 24.54   |
| Nisha Thakur (2019)      | 1.10 (0.78, 1.55) | 23.72   |
| Subtotal ($I^2 = 74.5\%, p = 0.048$) | 0.86 (0.55, 1.36) | 48.26   |

| $P > 0.05$               |               |          |
| Shizhi Wang (2019)       | 1.03 (0.87, 1.21) | 27.15   |
| Shruti Srivastava (2017) | 0.43 (0.31, 0.58) | 24.59   |
| Subtotal ($I^2 = 95.9\%, p = 0.000$) | 0.67 (0.28, 1.58) | 51.74   |
| Overall ($I^2 = 89.4\%, p = 0.000$) | 0.76 (0.50, 1.15) | 100.00  |

NOTE: Weights are from random effects analysis

(b)

Figure 4: (a) Forest plot of allele comparison of miRNA-499 rs3746444 for overall comparison (G versus A). (b) Forest plot of allele comparison of miRNA-499 rs3746444 for overall comparison (G versus A).
exist, because only published studies were included in this meta-analysis, and results that are not statistically significant usually have less chance of publication.

5. Conclusion

In summary, our results suggested that miR-124 rs531564 polymorphism is significantly associated with increased risk of cervical cancer, while miR-146a rs2910164, miR-196a2 rs11614913, miR-499 rs3746444, and miR-218 rs2292832 polymorphisms may not be associated with the susceptibility of cervical cancer. However, since few studies have been included, there is not enough data to fully confirm the association between cervical cancer and the five miRNA polymorphisms, so the results should be interpreted with caution.
Data Availability

The data supporting this meta-analysis are from previously reported studies which have been cited.

Conflicts of Interest

The authors declare that there are no conflicts of interest associated with the manuscript.

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Supplementary Materials

Table S1: scale for quality assessment. (Supplementary Materials)

References

[1] S. Meng, X. Fan, J. Zhang, R. An, and S. Li, “GJA1 expression and its prognostic value in cervical cancer,” BioMed Research International, vol. 2020, 10 pages, 2020.

[2] S. Srivastava, S. Singh, N. Fatima, B. Mittal, and A. N. Srivastava, “Pre-microRNA gene polymorphisms and risk of cervical squamous cell carcinoma,” Journal of Clinical and Diagnostic Research, vol. 11, no. 9, p. 4, 2017.

[3] A. C. de Freitas, A. P. Gurgel, B. S. Chagas, E. C. Coimbra, and C. M. do Amaral, “Susceptibility to cervical cancer: an overview,” Gynecologic Oncology, vol. 126, no. 2, pp. 304–311, 2012.

[4] J. Fang, Y. Li, J. Zhang et al., “Correlation between polymorphisms in microRNA-regulated genes and cervical cancer susceptibility in a Xinjiang Uygur population,” Oncotarget, vol. 8, no. 19, pp. 31758–31764, 2017.

[5] X. Zeng, Y. Zhang, T. Yue et al., “MicroRNA-361-5p inhibits tumorogenesis and the EMT of HCC by targeting Twist1,” BioMed Research International, vol. 2020, 14 pages, 2020.

[6] L. C. Yin, G. Xiao, R. Zhou et al., “MicroRNA-361-5p inhibits tumorigenesis and the EMT of HCC by targeting Twist1,” BioMed Research International, vol. 2020, 14 pages, 2020.

[7] K. Li, H. Tie, N. Hu et al., “Association of two polymorphisms rs2910164 in miRNA-146a and rs3746444 in miRNA-499 with rheumatoid arthritis: a meta-analysis,” Human Immunology, vol. 75, no. 7, pp. 31758–31764, 2017.

[8] M. A. Saunders, H. Liang, and W. H. Li, “Human polymorphism at microRNAs and microRNA target sites,” Proceedings of the National Academy of Sciences, vol. 104, no. 9, pp. 3300–3305, 2007.

[9] G. A. Calin and C. M. Croce, “MicroRNA-cancer connection: the beginning of a new tale,” Cancer Research, vol. 66, no. 15, pp. 7390–7394, 2006.

[10] L. Zhang, J. Huang, N. Yang et al., “MicroRNAs exhibit high frequency genomic alterations in human cancer,” Proceedings of the National Academy of Sciences, vol. 103, no. 24, pp. 9136–9141, 2006.

[11] T.-Y. Shi, X.-J. Chen, M.-L. Zhu et al., “A pri-miR-218 variant and risk of cervical carcinoma in Chinese women,” BMC Cancer, vol. 13, no. 1, p. 8, 2013.

[12] B. Zhou, K. Wang, Y. Wang et al., “Common genetic polymorphisms in pre-microRNAs and risk of cervical squamous cell carcinoma,” Molecular Carcinogenesis, vol. 50, no. 7, pp. 499–505, 2011.

[13] X. Xiong, J. Cheng, X. Liu, S. Tang, and X. Luo, “Correlation analysis between miR-124 rs351564 polymorphisms and susceptibility to cervical cancer,” Journal of Southern Medical University, vol. 34, no. 2, pp. 210–213, 2014.

[14] S. Wang, H. Zhu, B. Ding et al., “Genetic variants in microRNAs are associated with cervical cancer risk,” Mutagenesis, vol. 34, no. 2, pp. 127–133, 2019.

[15] N. Thakur, P. Singhal, R. Mehrotra, and M. Bharadwaj, “Impacts of single nucleotide polymorphisms in three microRNAs (miR-146a, miR-196a2 and miR-499) on the susceptibility to cervical cancer among Indian women,” Boscience Reports, vol. 39, no. 4, pp. 127–133, 2019.

[16] H. Wu and J. Zhang, “miR-124 rs531564 polymorphism influences genetic susceptibility to cervical cancer,” International Journal of Clinical and Experimental Medicine, vol. 7, no. 12, pp. 5847–5851, 2014.

[17] L. Chuanyin, W. Xiaona, Y. Zhiling et al., “The association between polymorphisms in microRNA genes and cervical cancer in a Chinese Han population,” Oncotarget, vol. 8, no. 50, pp. 87914–87927, 2017.

[18] C. Yue, M. Wang, B. Ding et al., “Polymorphism of the pre-miR-146a is associated with risk of cervical cancer in a Chinese population,” Gynecologic Oncology, vol. 122, no. 1, pp. 33–37, 2011.

[19] Z. Yan, Z. Zhou, C. Li et al., “Polymorphisms in microRNA genes play roles in the initiation and development of cervical cancer,” Journal of Cancer, vol. 10, no. 20, pp. 4747–4753, 2019.

[20] X. Zhou, X. Chen, L. Hu et al., “Polymorphisms involved in the miR-218-LAMB3 pathway and susceptibility of cervical cancer, a case-control study in Chinese women,” Gynecologic Oncology, vol. 127, no. 2, pp. 287–290, 2010.

[21] G. Chen, M. Zhang, J. Zhu et al., “Common genetic variants in pre-miRNAs are associated with cervical cancer susceptibility in southern Chinese women,” Journal of Cancer, vol. 11, no. 8, pp. 2138–2138, 2020.

[22] J. Guo, M. Jin, M. Zhang, and K. Chen, “A genetic variant in miR-196a2 increased digestive system cancer risks: a meta-analysis of 15 case-control studies,” PLoS ONE, vol. 7, no. 1, p. 7, 2012.

[23] D. Moher, L. Liberati, J. Tetzlaff, D. G. Altman, and PRISMA Group, “Preferred reporting items for systematic reviews and meta-analyses: the PRISMA statement,” Annals of Internal Medicine, vol. 151, no. 4, pp. 264–9, W64, 2009.

[24] J. L. Vanoverschelde, A. R. Robert, A. Gerbaux, X. Michel, C. Hanet, and W. Wijns, “Noninvasive estimation of pulmonary arterial wedge pressure with Doppler transmural flow velocity pattern in patients with known heart disease,” The American Journal of Cardiology, vol. 75, no. 5, pp. 383–389, 1995.

[25] J. Lau, J. P. Ioannidis, and C. H. Schmid, “Quantitative synthesis in systematic reviews,” Annals of Internal Medicine, vol. 127, no. 9, pp. 820–826, 1997.

[26] N. Mantel and W. Haenszel, “Statistical aspects of the analysis of data from retrospective studies of disease,” Journal of the National Cancer Institute, vol. 22, no. 4, pp. 719–748, 1959.
[27] R. DerSimonian and N. Laird, “Meta-analysis in clinical trials revisited,” *Contemporary Clinical Trials*, vol. 45, pp. 139–145, 2015.

[28] Z. M. Zheng and X. Wang, “Regulation of cellular miRNA expression by human papillomaviruses,” *Biochimica et Biophysica Acta*, vol. 1809, no. 11-12, pp. 668–677, 2011.

[29] K. Jazdzewski, E. L. Murray, K. Franssila, B. Jarzab, D. R. Schoenberg, and A. de la Chapelle, “Common SNP in pre-miR-146a decreases mature miR expression and predisposes to papillary thyroid carcinoma,” *Proceedings of the National Academy of Sciences*, vol. 105, no. 20, pp. 7269–7274, 2008.

[30] J. Shen, C. B. Ambrosone, R. A. DiCioccio, K. Odunsi, S. B. Lele, and H. Zhao, “A functional polymorphism in the miR-146a gene and age of familial breast/ovarian cancer diagnosis,” *Carcinogenesis*, vol. 29, no. 10, pp. 1963–1966, 2008.

[31] Y. Zeng and B. R. Cullen, “Sequence requirements for microRNA processing and function in human cells,” *RNA*, vol. 9, no. 1, pp. 112–123, 2003.

[32] T. Ito-Kureha, N. Koshikawa, M. Yamamoto et al., “Tropomodulin 1 expression driven by NF-κB enhances breast cancer growth,” *Cancer Research*, vol. 75, no. 1, pp. 62–72, 2015.

[33] I. Kastrati, E. Canestrari, and J. Frasor, “PHLDA1 expression is controlled by an estrogen receptor-NF-κB-miR-181 regulatory loop and is essential for formation of ER+ mammospheres,” *Oncogene*, vol. 34, no. 18, pp. 2309–2316, 2015.

[34] S. M. Wilting, R. A. van Boerdonk, F. E. Henken et al., “Methylation-mediated silencing and tumour suppressive function of hsa-miR-124 in cervical cancer,” *Molecular Cancer*, vol. 9, no. 1, p. 167, 2010.

[35] R. Garzon, G. A. Calin, and C. M. Croce, “MicroRNAs in cancer,” *Annual Review of Medicine*, vol. 2009, pp. 199–227, 2009.

[36] X. Wang, S. Tang, S. Y. le et al., “Aberrant expression of oncogenic and tumor-suppressive microRNAs in cervical cancer is required for cancer cell growth,” *PLoS One*, vol. 3, no. 7, p. e2557, 2008.

[37] N. Yamamoto, T. Kinoshita, N. Nohata et al., “Tumor suppressive microRNA-218 inhibits cancer cell migration and invasion by targeting focal adhesion pathways in cervical squamous cell carcinoma,” *International Journal of Oncology*, vol. 42, no. 5, pp. 1523–1532, 2013.

[38] A. Jemal, F. Bray, M. M. Center, J. Ferlay, E. Ward, and D. Forman, “Global cancer statistics,” *CA: a Cancer Journal for Clinicians*, vol. 61, no. 2, pp. 69–90, 2011.

[39] I. Torruella-Loran, H. Laayouni, B. Dobon et al., “MicroRNA genetic variation: from population analysis to functional implications of three allele variants associated with cancer,” *Human Mutation*, vol. 37, no. 10, pp. 1060–1073, 2016.

[40] A. E. Hoffman, T. Zheng, C. Yi et al., “MicroRNA miR-196a2 and breast cancer: a genetic and epigenetic association study and functional analysis,” *Cancer Research*, vol. 69, no. 14, pp. 5970–5977, 2009.

[41] P. Qi, T. H. Dou, L. Geng et al., “Association of a variant in MIR 196A2 with susceptibility to hepatocellular carcinoma in male Chinese patients with chronic hepatitis B virus infection,” *Human Immunology*, vol. 71, no. 6, pp. 621–626, 2010.

[42] K. Tano, R. Mizuno, T. Okada et al., “MALAT-1 enhances cell motility of lung adenocarcinoma cells by influencing the expression of motility-related genes,” *FEBS Letters*, vol. 584, no. 22, pp. 4575–4580, 2010.