Understory effect on tree and cork growth in cork oak woodlands

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Abstract

Aim of study: Cork oak is one of the main forest tree species in Portugal that typically occurs in montado, where operational practices oriented to the tree, crop or animal management may influence several of the ecosystem components. This study aimed at contributing to fulfill the lack of knowledge on the effect of these practices on the cork and wood growth, by comparing the wood diameter growth and the annual cork increment under two different understory management options.

Material and methods: An experimental trial implemented on an uneven-aged cork oak pure stand during a cork rotation period of 9 years, was established with the specific goal of comparing understory management options: a yellow lupine pasture versus spontaneous vegetation. Cork samples were taken at the beginning and end of the period and were used to measure cork thickness and annual cork rings. The differences between treatments were assessed performing a non-parametric test and a more robust approach using linear mixed model. Precipitation and treatment levels were jointly considered on the analysis.

Main results: A slight effect was found on the cork thickness regarding the treatment with lupine application. However, no distinct effect was found, regarding wood and the annual cork increment pattern. Additionally, annual cork ring width showed a positive correlation with precipitation and a negative correlation with ring age.

Research highlights: The results of this study indicate no distinct pattern regarding the annual cork and wood increment when comparing the understory effect of yellow lupine pasture versus spontaneous vegetation.

Additional keywords: Quercus suber, cork thickness; cork ring; lupine; shrubs; linear mixed model.

Introduction

Cork oak (Quercus suber) is a Mediterranean species covering a worldwide area of 2,139,942 ha (APCOR, 2016), from which the tree bark (cork) is extracted and used as raw material of an industry that is responsible for a total value of exportations of 1,430 million € (APCOR, 2016). In Portugal, a country responsible for supplying 49.6% of the world cork production (APCOR, 2016), this tree plays a key role mostly in the southern rural areas. It is the main forest species present in the traditional silvopastoral system called montado, which is characterized by tree densities around 66 trees/ha, according to the last published official data (AFN, 2010).

In montado, management practices are oriented not only for cork extraction purposes, but also for grazing or crop production, which may have influence on several ecosystem components (e.g. Pinto-Correia et al., 2011; Paulo et al., 2016a). Under a silvopastoral management, it is a common practice to periodically install a legume rich pasture in the understory, such as Lupinus luteus, with the expectation of improving soil fertility and decrease the need of external input for animal forrage (e.g. Callaway, 1995; Teixeira et al., 2011). When animals are not present, the spontaneous
vegetation is periodically mechanically removed, reducing the fuel component to avoid fire hazard. These mechanization processes, have been referred to have a negative impact on soil compaction and affect the tree roots development contributing to the tree decline (e.g. Dinis et al., 2015).

It is expected that interspecific competition between the trees and the understory influences tree growth patterns, along with intraspecific competition and stand structure, depending on the species, environment and management conditions. In Mediterranean cork oak stands, Sánchez-González et al. (2006) found out that tree diameter growth is negatively influenced by increasing stand density, however Paulo et al. (2016b) did not find the same pattern. Regarding the competition between trees and the understory species in the neighbourhood, diverse relationships should be considered. The presence of understory may be positive for some soil functions as it contributes to the nutrient content (Moreno & Obrador, 2007), and may influence tree natural regeneration (Pulido & Díaz, 2005). The Cistus species is a common spontaneous species in the cork oak woodland understory, that is reported to promote different soil C/N ratios (Gómez-Rey et al., 2013). According to Correia et al. (2014), the Cistus salviifolius species success relates to a high shadow tolerance and a fast water uptake in short precipitation events. Caldeira et al. (2015) experimental results with Cistus ladanifer indicate that this species is more tolerant to low soil water potential and dry soil conditions than cork oak trees.

Several studies focus on the prediction of cork growth evolution (e.g. Sánchez-González et al., 2007; Almeida et al., 2010; Paulo et al., 2016b); however, few studies are available when focusing the effect of competition between trees and understory species. This is even more evident when considering the cork oak species, where both the effect on tree wood growth and on cork growth are of interest. Caritat et al. (1999) indicated that there is no significant effect of shrubs presence on cork tree radial growth, although apical elongation was higher in the absence of shrubs. Similar results were found by Martín et al. (2015) for the holm oak species, since no significant effects were found on tree growth when comparing understory and soil management practices. Thus, no results were found when considering cork and wood differentiation growth.

To increase our understanding on the effects of the understory on the cork and tree growth, this study focuses not only in wood diameter growth, but also in the cork annual growth. The research addresses the following question: do trees have a similar increment pattern regarding wood and cork, when growing under two different understory management options?

Material and methods

Trial description

The experimental trial was located in Portugal Center, near the Montargil village (39°3.242’N, 8°10.588’W) and implemented in 2003, in an uneven-aged mature cork oak stand. The stand was characterized by an average of 101 trees/ha, where the major percentage of trees have already been debarked. There were two different cork rotation cycles within the stand: from 2003 to 2012 and from 2006 to 2015, the years in which cork stripping was performed. The understory layer, composed by spontaneous vegetation dominated by the Cistus salvifolius, alongside with sparsely distributed Rosmarinus officinalis and Ulex airesis, was characterized by a phytomass of 0.35 kg/m² at the age of 4 years. From 1994 to 2003 the understory was mechanically removed with an interval of 3 to 4 years. Accessing the digital cartography available at the Portuguese Agency for the Environment website (http://sniamb.apambiente.pt/webatlas/) soils were classified as Podzols, according to the IUSS Working Group WRB (2006). Plot determination of pH and organic matter percentage (OM) was done in 2009 (Table 1).

The trial focused on the comparison of two different understory management options, from now on designated as treatments, implemented as a complete randomized block design. Each block consisted of two quadrangular plots with 2 ha, allowing the delimitation of a 20 m border to ensure no impact of non-treated areas on the trees used for the experiment. The first treatment consisted in the maintenance of spontaneous understory vegetation during the complete cork debarking rotation period (NUR, no understory removal). The second treatment consisted in the periodical removal of the understory with incorporation of organic matter into the soil, followed by the seeding of Lupinus luteus (RUL, understory removal with lupine seeding). This treatment was repeated along the cork debarking rotation period in 2003, 2007 and 2009. After each application, visual assessment on the success of the lupine germination rate was made in the field.

Tree measurements

Measurements of tree diameter at breast height after debarking (du) were carried out at each debarking year, in 2003 and 2012, for all the trees inside the plots after being debarked. The wood diameter increment (idu) was computed as the difference between the diameter at breast height under cork (du) from the debarked trees measured in 2012 and 2003 (\( n_{\text{block1}} = 302; n_{\text{block2}} = 229 \)). The stand characteristics were computed at the beginning and
at the end of the period, namely: number of trees per hectare (N); the quadratic diameter at breast height under cork (dug) and crown cover percentage (Cc) (Table 1).

Cork samples collected

In 2003 and 2012, cork samples with approximately 20 cm × 20 cm, were taken at breast height from the debarked trees. Samples from 2003 cover the cork growth period between 1994 and 2003, when the understory was mechanically removed with an interval of 3 to 4 years. Samples from 2012 cover the cork growth period between 2003 and 2012, when the understory management was differentiated by the treatments implementation. All samples were boiled during 1 hour in water at 100ºC and atmospheric pressure, and left to air-dry in well ventilated conditions until equilibrium. The aim is to decrease internal tensions, caused by the cellular corrugation during cork growth, that are particularly important in the radial direction where cork thickness is measured (Pereira, 2007). Cork thickness before and after boiling was measured on each cork sample using a digital caliper (n_{block1} = 94; n_{block2} = 57). Annual cork growth rings, a total of 8 complete rings, were measured, after been visually identified on at least two positions, using the image analysis software ImageJ (Ferreira & Rasband, 2010; Schneider et al., 2012). This procedure was carried out only in the trees having the cork samples in both debarking years where the rings were unmistakable marked (n_{block1} = 75; n_{block2} = 51).

Climate data

According to the known relationship between annual cork growth and precipitation (e.g. Caritat et al., 2000; Paulo et al., 2016b), values of monthly precipitation for the 1994 to 2012 period were gathered from the nearest meteorological station in Montargil, available at the network SNIRH - Sistema Nacional de Informação de Recursos Hídricos (www.snirh.pt).

Data analysis

As a first step, the cork thickness, the annual cork ring width and the wood diameter increment measurements were used for a preliminary graphical analysis. The empirical distributions of the cork thickness and of the annual cork ring width were compared among treatments and separately for each block. The hypotheses to test were: 1) the distributions were not significantly different in 2003; 2) the distributions became significantly different in 2012, after the treatments application. The hypotheses were tested using the non-parametric statistical test of Kruskal-Wallis (McDonald, 2014), separately for the cork samples of 2003 (previous to the trial establishment) and for the cork samples of 2012 (subject to the trial treatments: RUL and NUR). The annual cork ring width was tested for each ring year. These analyses were evaluated (α = 0.05), with the PROC NPAR1WAY procedure of the SAS 9.3 (SAS Inst., 2011).

Modelling approach

Although the trial was established following a complete randomized block design, due to the nested structure of the data, trees inside plots and plots inside blocks, the data analysis was carried out using a mixed modelling approach. This parametric approach is particularly suitable for data referring to growth curves since it considers not only the nested structure of the data, but also the correlation from the repeated measures taken on the same individuals (e.g. Pinheiro

Table 1. Plot characterization by block, inventory year and treatment, where RUL is the treatment with understory removal and lupine pasture and NUR is the treatment with spontaneous understory vegetation maintenance.

| Block | Treatment | Year | dug | N  | Cc (%) | M (%) | pH | OM (%) |
|-------|-----------|------|-----|----|--------|-------|-----|--------|
| 1     | RUL       | 2003 | 27.5| 132| 55     | 1     | 5.5 | 1.82   |
|       |           | 2012 | 28.6| 131| 58     |       |     |        |
|       | NUR       | 2003 | 27.9| 127| 53     | 12    | 6.1 | 1.47   |
|       |           | 2012 | 28.1| 114| 49     |       |     |        |
| 2     | RUL       | 2003 | 36.1| 86 | 53     | 22    | 5.6 | 1.45   |
|       |           | 2012 | 37.6| 71 | 46     |       |     |        |
|       | NUR       | 2003 | 33.6| 100| 55     | 11    | 5.7 | 1.37   |
|       |           | 2012 | 35.7| 91 | 56     |       |     |        |

dug, quadratic diameter at breast height under cork (cm); N, number of trees/ha; Cc, crown cover percentage; M, percentage of dead trees; OM, organic matter percentage.
& Bates, 2000), specifically the eight complete rings of each cork sample. For the cork annual growth, a linear mixed model was developed using the samples from 2012. The development of this model was preceded by the selection of the precipitation variable that was best related to the variable cork annual growth. The potential variables were retrieved from the list presented by Caritat et al. (2000) and were computed between October 1st of the year before the growth period and September 30th of the growth period year. This process was performed by fitting the following fixed base model:

\[ rw_i = \beta_0 + \beta_1 t_i + \beta_2 P_i + \epsilon_i, \]  

[1]

where \( rw \) is annual cork ring width (mm) of the tree \( i \); \( t \) is the cork age; \( P \) is the volume of precipitation for a specific period; \( \beta_0, \beta_1, \beta_2 \) are the fixed effects parameters and \( \epsilon \) is the error term.

The selected precipitation variable was the one that resulted in the lower value of the mean square error (MSE) and Akaike information criterion (AIC). The fitting of the linear mixed model was then carried out by adding to the base model (Eq. [1]), a dummy variable concerning the treatment with lupine (RUL), a variable regarding the interaction between the cork age and the treatment, and the plot’s random effects. Additionally, a linear mixed model was developed for the tree wood diameter increment that included variables of the initial tree diameter, the dummy variable concerning the treatment with lupine (RUL), and the plot’s random effects. The tree wood diameter increment, was assessed between the two consecutive cork extractions, performing 9 years growth. The dataset considered for fitting included all the debarked trees measured in 2003 and 2012.

All models were fitted using the procedure PROC MIXED of the software SAS 9.3 (SAS Inst., 2011). The variance components were used for the covariance structure. The RANDOM statement was applied to specify the random effects associated to the plot levels. The random effect parameter was tested in all the fixed parameters, and the selection criteria for its inclusion in the model was the lowest AIC value. For each developed model, the significance of the parameters estimates were evaluated considering \( \alpha = 0.05 \).

Results

Treatment effect in cork growth

There was a clear decrease in the cork thickness measured after boiling, assessing the 2003 and 2012 subset samples observed in both blocks (Fig. 1). The Krushal-Wallis (KW) results performed for the cork thickness empirical distribution for the 2003 samples presented no significant difference between treatments in both blocks (\( p=0.2246 \) in block 1; \( p=0.5388 \) in block 2). This confirms that at the beginning of the trial the plots were similar. The KW test results obtained for the cork thickness, using the 2012 samples, indicate a significant difference between NUR and RUL treatments in block 1 (\( p=0.0007 \), mean RUL = 27.62, mean NUR = 23.15), but no significant difference in block 2 (\( p=0.2995 \), mean RUL = 24.45, mean NUR = 26.37).

When visualizing the annual cork ring width of each treatment (Fig. 2), no similar pattern was found in either blocks, for each subset of cork samples, 2003 and 2012. Regarding the KW test results for the annual cork ring width using the 2003 samples, no significant difference was found between the two blocks (results not presented), confirming once more that both blocks were similar before the trial establishment. While, the KW test results for the annual cork ring mean width using the 2012 samples showed significant differences between NUR and RUL treatments on block 1 for the years: 2004, 2005 and 2009, all corresponding to the year of lupine seeding or the following year (Table 2). The only exception in block 1 was the year 2007 (Table 2), however the lupine seeding in this year was applied under unsuitable precipitation and soil humidity conditions, resulting in a low germination rate of the lupine during the following year. The KW test results or the annual cork ring width of block 2 were not the same when comparing NUR and RUL treatments, only in 2010 the annual cork ring distribution was different between the treatments (Table 2). Regarding the annual cork ring width model, the precipitation variable included was the one selected according to the MSE and AIC lower values: the annual precipitation for the period between October 1st of the year before the growth period and September 30th of the growth period year (Table 3). For this model, the random effect was more significant when added only to the intercept parameter, since it was not possible to obtain convergence with more than one random effect. Thus, the full random effect model defined was:

\[ rw_{2012ij} = (\beta_0 + \mu_i) + \beta_1 t + \beta_2 P_i + \beta_3 RUL_j + \beta_4 \left(t \times RUL_j\right) + \epsilon_{ij}, \]  

[2]

where \( rw \) is the annual cork ring width of the tree \( i \); \( t \) is the cork age; \( P \) is the volume of precipitation for a given period; \( RUL \) is the variable regarding the treatment with lupine; \( \beta_0, \beta_1, \beta_2, \beta_3, \beta_4 \) are the fixed effects parameters; \( \mu_i \) is the random effect associated to plot \( j \) and \( \epsilon_{ij} \) is the error term.
For the full fitted model (Eq. 2), the parameters estimates for \( P \) and \( t \) were statistical significant and showing a positive correlation with precipitation and a negative correlation with ring age (Table 4). The treatment with lupine (RUL) was not significant, while the interaction between the cork age and the RUL treatment was significant. However, when excluding the RUL variable, this interaction was also not statistical significant. The random effects term were not significantly different from zero for any of the plots. Thus, the final random effect model only included precipitation and cork ring age (Table 4),
which is an indication that the treatments are not different concerning their influence in annual cork ring increment.

### Treatment effect in tree wood diameter increment

The wood mean diameter increment plot with diameter classes showed no similar pattern on the treatments by block (Fig. 3). For the wood mean diameter increment model, the random effect was also more significant when added to the intercept parameter. The full random effect model defined was:

\[
\text{idu}_{ij} = (\beta_0 + \mu_0) + \beta_1 \text{du}_{2003i} + \beta_2 \text{RUL}_j + \epsilon_{ij},
\]

where \text{idu}_i is the wood diameter increment (mm) for the debarking period of tree \(i\); \text{du}_{2003i} is the tree diameter under cork measured in 2003; \text{RUL}_j is the variable regarding the treatment with lupine; \(\beta_0, \beta_1, \beta_2\) are the fixed parameters; \(\mu_0\) is the random effect associated to plot \(j\) and \(\epsilon_{ij}\) is the error term.

For the full fitted model (Eq. [3]), the tree diameter under cork measured in 2003 was statistical significant showing a positive correlation. However, the \text{RUL} treatment was not significant, but also the intercept parameter was not significantly different from zero (Table 5). Thus, the final random effect model only included the tree diameter under cork (Table 5), indicating also that the treatments are not different concerning their influence in wood diameter increment.

### Discussion

This work was the first specifically dedicated to the purpose of evaluating the influence of different understory management options on cork annual growth and tree wood growth, a research question already referred by Oliveira & Costa (2012) and Paulo et al. (2016b). For this purpose, the trial allowed to obtain two subsets of cork samples that were used for the analysis.
of cork thickness and cork ring width under two distinct understory management regimes. This is one of the study main features, since it allowed to characterize the annual cork growth rates and distributions previous to the treatments application, demonstrating that not only management was similar along the four different plots, but also that the cork annual growth presented equal distributions among the four plots.

Cork thickness decreased from the 2003 to the 2012 samples, irrespective of the treatment, which may be the result of a decrease of total precipitation within the second debarking period, and in line with the existing literature on the relationship between annual cork growth and climate conditions (e.g. Caritat et al., 2000). The fitted model for the annual cork ring width showed a positive correlation with precipitation, as shown in several other studies (e.g. Paulo et al., 2016; Oliveira et al., 2016). Regarding the cork ring age, a negative correlation was obtained, as expected from known studies (e.g. Costa et al., 2002; Pereira, 2007; Oliveira et al., 2016), that showed a decrease of the annual cork width trend along the cork debarking period. The inclusion of precipitation and cork age variables in the fitted baseline model for annual cork growth, allowed to isolate the treatment effect, as well as, the plot and block random effects.

No difference was found between the two understory treatments regarding their influence in annual cork width or wood diameter increment, in line with Caritat et al. (1999) related study, in which the results for a sample of 10 trees by treatment did not show significant differences between the treatments. Nevertheless, it was also observed that plots in both blocks did not present a similar response to the treatments. One block presented a small positive effect of the understory removal treatment, with an increase of less than 1 mm on the mean annual cork growth, when the lupine was applied or in the following year, whereas in the other block a small effect was detected only in one year. Looking at the annual response of the tree to the treatments, it was observed that although the lupine installation may favour cork increment one or two years after application, this effect could also be inexistent in years when unfavourable conditions prevailed. The increments lower than 1 mm are not expected to have an impact in the cork price and the resulting farm income (Paulo & Tomé, 2017). This varying effect suggests that site characteristics, such as soil and stand structure, may influence the impact of different management options, but that these may also be related to annual climate conditions (Sánchez-González et al., 2007).

Mechanical operations such as the ones carried out for understory removal, may affect the tree roots development, contribute to the tree decline (e.g. David et al., 2013), or limit cork oak regeneration (e.g. Arosa et al., 2017). These operations being frequently performed may slow the full recovery of understory composition and structure as evidence in Santana et al. (2011). We recommend that management plans should be frequently reviewed, at the stand scale, regarding the conditions, the management goals such as the cork and cattle production, the tree regeneration, the fire hazard reduction and the biodiversity conservation, following an adaptive management strategy (e.g. Aronson et al., 2009).

**Conclusion**

The data used for this study, were collected from a trial established with the specific goal to contribute for understanding the effects of the understory on the cork and tree growth. The study outcome indicates no effect, for this site, regarding cork or wood increment pattern, when comparing the understory effect of yellow lupine pasture versus spontaneous vegetation. The cork thickness variability between trees and the individual tree response to annual climate conditions is more determinant to the final cork thickness than the management alternatives considered in this research. Nevertheless, the differences found across the two blocks suggest that site characteristics should be explored in further research. The follow-up monitoring of this trial and the establishment of

| Precipitation variable (defined by the period considered) | Model | MSE | AIC |
|----------------------------------------------------------|-------|-----|-----|
| March<sub>t</sub> to June<sub>t</sub>                   |       | 1.9712 | 5422.8 |
| March<sub>t</sub> to May<sub>t</sub>                   |       | 2.5306 | 5806.0 |
| March<sub>t</sub> to September<sub>t</sub>             |       | 1.8495 | 5326.2 |
| January<sub>t</sub> to June<sub>t</sub>                |       | 2.0706 | 5500.0 |
| January<sub>t</sub> to September<sub>t</sub>           |       | 2.0444 | 5480.5 |
| June<sub>t</sub> to September<sub>t</sub>              |       | 2.3202 | 5673.4 |
| October<sub>t</sub> to September<sub>t</sub>           |       | 1.7569 | 5249.6 |
| October<sub>t</sub> to June<sub>t</sub>                |       | 1.8590 | 5336.1 |
| October<sub>t</sub> to December<sub>t</sub>           |       | 2.1945 | 5589.8 |

<sub>t</sub>: year of the growth period
Table 4. Parameters estimates for the annual cork ring width ($rw$) random fitted models from Eq. [2] ($\text{AIC}_{\text{Full}} = 2184; \text{AIC}_{\text{Final}} = 2182$).

|                         | Full model | Final model |
|-------------------------|------------|-------------|
|                         | Estimate   | $p$-value   | Lower | Upper     | Estimate   | $p$-value   | Lower | Upper     |
| Fixed                  |            |             |       |           |            |             |       |           |
| $\beta_0$              | 2.1913     | 0.0119      | 1.1518| 3.2308    | 2.3635     | 0.0007      | 1.8358| 2.8911    |
| $\beta_{tc}$           | -0.1016    | <0.0001     | -0.1434| -0.0598  | -0.1290    | <0.0001     | -0.1613| -0.0967  |
| $\beta_{Pr}$           | 0.0024     | <0.0001     | 0.0019| 0.0028    | 0.0024     | <0.0001     | 0.0019| 0.0028    |
| $\beta_{RUL}$          | 0.3915     | 0.2306      | -0.2491| 1.0321   |              |             |       |           |
| $\beta_{tc \times RUL}$| -0.0625    | 0.0443      | -0.1235| -0.0016  |              |             |       |           |
| Random                 |            |             |       |           |            |             |       |           |
| $\mu_{b_{11}}$         | 0.1229     | 0.5316      | -0.2627| 0.5086   | 0.1403     | 0.2668      | -0.1076| 0.3882   |
| $\mu_{b_{12}}$         | -0.1229    | 0.5316      | -0.5086| 0.2627   | -0.0955    | 0.4402      | -0.3382| 0.1472   |
| $\mu_{b_{21}}$         | -0.227     | 0.2469      | -0.6115| 0.1576   | -0.2424    | 0.0447      | -0.4790| -0.0058  |
| $\mu_{b_{22}}$         | 0.227      | 0.2469      | -0.1576| 0.6115   | 0.1975     | 0.1220      | -0.0529| 0.4479   |
| Cov. Parm              |            |             |       |           |            |             |       |           |
| intercept $\mu_{b_0}$  | 0.0718     | –           | 0.01827| 4.6625   | 0.04713    | –           | 0.01389| 1.0591   |
| Residual               | 0.9566     | –           | 0.8673| 1.0604   | 0.9604     | –           | 0.8708| 1.0646   |

Table 5. Parameters estimates for the wood diameter increment ($idu$) linear mixed fitted models from Eq. [3] ($\text{AIC}_{\text{Full}} = 1884; \text{AIC}_{\text{Final}} = 1892$).

|                         | Full model | Final model |
|-------------------------|------------|-------------|
|                         | Estimate   | $p$-value   | Lower | Upper     | Estimate   | $p$-value   | Lower | Upper     |
| Fixed                  |            |             |       |           |            |             |       |           |
| $\beta_0$              | 1.1489     | 0.1476      | -0.9956| 3.2933   |              |             |       |           |
| $\beta_{tc_{200}}$     | 0.0207     | 0.0185      | 0.0035| 0.0379   | 0.0287     | 0.007      | 0.01   | 0.0438   |
| $\beta_{RUL}$          | 0.0274     | 0.9631      | -1.1358| 1.1906   |              |             |       |           |
| Random                 |            |             |       |           |            |             |       |           |
| $\mu_{b_{11}}$         | -0.3198    | 0.4471      | -1.1456| 0.5060   | 0.6044     | 0.016      | 0.1131| 1.0957   |
| $\mu_{b_{12}}$         | 0.3198     | 0.4471      | -0.5060| 1.1456   | 1.2090     | <0.0001    | 0.6107| 1.8074   |
| $\mu_{b_{21}}$         | -0.4662    | 0.2643      | -1.2857| 0.3534   | 0.4248     | 0.0917     | -0.0691| 0.9186   |
| $\mu_{b_{22}}$         | 0.4662     | 0.2643      | -0.3534| 1.2857   | 1.3574     | <0.0001    | 0.8178| 1.8970   |
| Cov. Parm              |            |             |       |           |            |             |       |           |
| intercept $\mu_{b_0}$  | 0.3352     | –           | 0.0872| 18.019   | 1.0354     | –           | 0.3004| 25.716   |
| Residual               | 1.9551     | –           | 1.7388| 2.2147   | 1.9580     | –           | 1.741| 2.2185   |
similar trials is needed in order to clarify the long-term tree response in consecutive cork debarking periods and in different environmental and structural stand characteristics.

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