Differential Cost-Sharing Undermines Treatment Adherence to Combination Therapy: Evidence from Diabetes Treatment

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ABSTRACT

Introduction: The objective of this study was to measure the influence of differences in out-of-pocket (OOP) costs for type 2 diabetes (T2D) medications on within-patient adherence behavior towards combination drug therapy regimens.

Methods: This was an observational, retrospective, paired sample study in patients with T2D using longitudinal pharmacy data from the 2009–2014 Medical Expenditure Panel Survey (MEPS) augmented with socio-demographic factors. We took a within-patient approach to minimize confounding effects by including patients who maintained the same number of diabetes drug classes over 2 years of MEPS. For each patient, we selected the most and least costly drug classes in the second year and examined their corresponding adherence behavior measured by medication possession ratio. The primary hypothesis tested the significance of the correlation between magnitude of the OOP cost difference and behavioral response in adherence.

Results: Analysis included 1189 patients representing over 4.2 million US residents with T2D. A significant negative correlation ($p < 0.001$) was observed between the differences of OOP costs and adherence to the most and least costly medications compared within patients. Reduction in adherence to the most costly medication was generally observed when the difference in OOP costs was greater than $33/month. A greater variability in adherence was observed when the cost difference exceeded $2.39/month as compared to other cost difference ranges ($p < 0.001$), indicative of choices being made.

Conclusions: As OOP costs increased, adherence variability increased initially until a cost threshold, beyond which the adherence to the more costly medication decreased. In addition to OOP cost, adherence was also influenced by type of medication and self-perception of health. Given the complex correlation between OOP costs and adherence to medication, we suggest a careful approach to cost-sharing in the current insurance drug design and relevant insurance policies.

Keywords: Adherence; Healthcare out-of-pocket costs; Diabetes combination therapy; Insurance design; Medical Expenditure Panel Survey (MEPS)
**Key Summary Points**

**Why carry out this study?**

Multiple therapy becomes increasingly important for patients with type 2 diabetes (T2D) and current insurance designs may make some patients less adherent to medications. Adherence to diabetes medications is low when out-of-pocket (OOP) costs and insurance co-payments are high. A 10% increase in patient cost-sharing may decrease adherence by 5.4–6.2% depending on medication class.

Since out-of-pocket (OOP) costs for the individual components of combination therapy may differ, it is not known whether OOP cost differences impact adherence to combination therapy. The study examined the correlation between magnitude of the OOP cost difference and behavioral response in adherence to combination therapy.

**What was learned from the study?**

Using the within-patient approach for analysis, we were able to minimize the impact of confounding factors on the relationship between OOP costs and adherence behavior, which is a common issue in cross-sectional studies.

Within the cohort, patients behave differently to cost differences to the medications they are taking. As OOP costs increased, the effect on adherence initially was more variable until a cost threshold beyond which the adherence to the more costly medication decreased.

Given the complex correlation between OOP costs and adherence to medication, we suggest a careful approach to cost-sharing in the current insurance drug design and relevant insurance policies.

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**INTRODUCTION**

As a chronic metabolic disorder, diabetes poses a significant economic burden on the United States (US) healthcare system. A portion of the high cost of diabetes medical expenditure was considered to be contributed by overconsumption of healthcare resources induced by the availability of health insurance, and thus has been viewed as an economic inefficiency that would lead to “welfare loss,” or decreased economic well-being [1, 2]. To control the cost, patients with diabetes are offered cost-sharing schemes in the form of deductibles, co-insurance, or co-payments [3]. However, the greater adoption of consumer-directed or high-deductible health plans has led to an unprecedented increase in initial out-of-pocket (OOP) costs [4, 5]. Cost-sharing schemes may be effective in reducing the economic inefficiency if patients are able to rationally weigh in the benefits and risks of the healthcare resources available to them, as is assumed by the traditional economic theory that was based on insurance-induced moral hazard theory [6, 7].

Breaking away from this traditional theory, some economists [8, 9] argue that imposing high cost-sharing may negatively impact patient health. Patients with serious illnesses may not have sufficient experience and knowledge to make appropriate medical choices and may opt for less effective therapy if essential care is not affordable. High cost-sharing could render healthcare unaffordable for people with low incomes. The Affordable Care Act (2010) offered free preventive services and provided insurance coverage to some people from the lower income group, thereby reducing their OOP spending. However, a majority of patients continue to experience a high burden of OOP and insurance premium expense [10]. Higher
cost-sharing may lead to delayed drug initiation and reduced adherence to medication, which contributes to poor health outcomes and ultimately higher healthcare resource utilization, defeating the purpose of cost-sharing [11–14].

As T2D is a progressive disease, maintenance of glycemic control with combination therapy often becomes necessary over time [15]. Combination therapy during the early treatment of T2D allows faster attainment of glycemic goals compared with sequential addition of medications that may result in primary and secondary failure [16–18]. In the tiered cost-sharing insurance design, patients often face different OOP costs for each of the medications within their combination therapy regimen. The difference in the OOP costs is likely to be even larger for those in high-deductible or consumer-driven insurance plans.

For patients with type 2 diabetes (T2D), poor adherence to therapy is particularly detrimental with increased difficulty in managing its downstream complications [19–21]. As adherence to medication may be significantly influenced by OOP costs [22–24], the association between adherence and OOP costs in patients with T2D has been widely studied but with several methodologic variations [25]. The results vary greatly across studies, ranging from no association to a significant inverse relationship [23]. A challenging aspect of studying OOP costs and adherence is potential confounding factors, which make it difficult to estimate the specific contribution of their relationship to adherence.

In this study, we analyzed data from patients utilizing combination therapy for T2D and examined the correlation between OOP costs and medication adherence. We study whether OOP cost differences impact adherence to their combination therapy. Rather than trying to capture the large number of measured and unmeasured interpatient confounding variables (much less integrate their collinearity), or to identify and successfully implement instrumental variables, our study design simplified the analysis in a novel way. By using paired samples with the concept of individuals serving as their own controls rather than a cross-sectional approach, we controlled any causal and confounding factors and reduced interpatient variability of measurable and non-measurable confounders.

METHODS

Sample Selection

We analyzed five panels (2009–2014) from the Medical Expenditure Panel Survey (MEPS) database [26] and selected individuals identified as having T2D and who utilized at least two antidiabetic drug medications for both years without intensifying or discontinuing the treatment, which would otherwise confound the analyses (see Supplementary Materials for details). We assumed: (a) classes of diabetes medications acquired in the first year of the MEPS panel (during at least two rounds) should have been acquired in the second year of the panel and (b) drugs in the same class with the same mechanism of action were considered interchangeable and drugs in different classes as complements (add-on therapy) [27]. The role of the first-year data was to select the combination therapy sample of patients. The second-year data were used for the analyses. For brevity, we refer to medication class as medication.

This study was conducted with secondary-use, publicly available, and de-identified data. MEPS has been reviewed and approved by the Westat IRB, established under a multi-project assurance (MPA M-1531) granted by the Office for Protection from Research Risks (OPRR). The project is reviewed, and the approval is renewed annually.

Measures of OOP Expenditures

The last purchase in the second year was used to calculate OOP spending per day supplied (see Supplementary Materials for details) [27]. The medication classes with highest and lowest OOP costs per day were selected for each patient and difference in OOP costs (ΔOOP) and difference in medication possession ratio (AMPDR) were calculated. As a result of the skewed distribution of ΔOOP, we defined ten categories, from low to
high ΔOOP. The ten deciles were further categorized as least OOP cost difference group (least ΔOOP), moderate OOP cost difference group (moderate ΔOOP), and large OOP cost difference group (large ΔOOP) with the cost difference cutoff determined post hoc on the basis of adherence patterns.

**Measure of Adherence**

We measured adherence using MPR, which was defined as the days' supply divided by 365 days in a year, subtracted by any nights spent in the hospital as an inpatient. Generally, quantity of medication dispensed (e.g., number of pills) was available in MEPS, but days' supply contained missing values. Therefore, for each National Drug Code (NDC), median days per unit from 2010 to 2014 was calculated to impute missing values [27]. All days' supplies were summed over their second MEPS year for all medications in the same class. If the summed days' supply was greater to or equal to 365 days (minus any nights spent in the hospital as an inpatient), the MPR was assigned a value of 1.0.

Two variables were derived using MPR. First, ΔMPR was defined as the difference between the MPR of the most costly OOP medication and that of the least costly OOP medication with a minimum and maximum ΔMPR of –1 and +1. Second, a categorical variable for adherence behavior was defined. When a conventional cutoff was used, a patient was considered adherent to a medication if his/her MPR was ≥80%. As adherence to two medications were examined at the same time for each patient, the adherence behavior variable had four types: (a) “adhere to both” had MPR ≥80% for both most and least costly medications; (b) “adhere to neither” had MPR < 80% for both medications; (c) “adhere to the most costly medication” had MPR ≥80% for only the most costly medication; and (d) “adhere to the least costly medication” had MPR ≥80% for only the least costly medication.

**Statistical Analysis**

To understand whether and how patient adherence varied by the OOP cost difference (ΔOOP) between the medications (the primary hypothesis), a univariate regression was performed with ΔMPR as the dependent variable and deciles of ΔOOP cost as the independent variable. Based on a post hoc examination of the results, a complex survey version of the Brown–Forsythe test [28] was performed to determine whether the variability of the ΔMPR among patients increased as the ΔOOP between medications increased. If the patients treated the two medications in their combination therapy as complements, ΔMPR would not be responsive to ΔOOP cost.

A logistic multinomial model was employed to further understand which factors influenced the patients’ adherence behavior to their combination therapy. The reference group was the group adherent to both most and least costly medications. For the selection of the independent variables, exploratory hypothesis-generating classification and regression tree (CART), as well as canonical discriminant analyses were first performed (further details and results in the Supplementary Materials).

Estimates were design-based using the strata, clusters, and weights of the complex survey design. Continuous variables and percentages were expressed as mean ± SE, except for scatter plots of OOP costs that utilized geometric means owing to skewness. Analysis were performed in SAS version 14.1 and JMP® version 13.1 (SAS Institute Inc., Cary, NC, USA).

**RESULTS**

**Demographics and Socioeconomic Characteristics**

A sample of 1189 patients were identified, representing 4,294,336 US patients with T2D on combination therapy (Table S1). Their average age was 62.4 ± 0.5 years (Table 1). Mean family income for first year was $51,845 ± $2078 (median $39,660). The average OOP cost for prescription medications in year 1 was
Table 1 Summary statistics of the patient characteristics, adherence (MPR), and out-of-pocket (OOP) costs among the overall selected population and within patient subgroups based on the difference in OOP between their most and least costly diabetes medication class

| Variables | Least ΔOOP | Moderate ΔOOP | High ΔOOP | Overall |
|-----------|------------|---------------|-----------|---------|
|           | Mean ± SE  | Mean ± SE     | Mean ± SE | Mean ± SE |
| n         | 488        | 506           | 195       | 1189    |
| Weighted n | 1,609,887  | 1,885,276     | 799,173   | 4,294,336 |
| Age, years | 63.55 ± 0.66 (63.26) | 61.06 ± 0.68 (61.05) | 63.39 ± 1.10 (62.93) | 62.43 ± 0.46 (62.15) |
| Sex (male), % | 45.91 ± 3.07 | 53.26 ± 2.71 | 53.80 ± 4.25 | 50.61 ± 1.87 |
| BMI, kg/m² | 32.60 ± 0.44 (31.12) | 32.60 ± 0.35 (32.34) | 32.91 ± 0.64 (31.90) | 32.66 ± 0.26 (31.91) |
| Age at T2D diagnosis, years | 51.33 ± 0.68 (50.97) | 48.83 ± 0.65 (49.12) | 50.66 ± 1.24 (49.76) | 50.11 ± 0.48 (49.55) |
| Family size, n | 2.27 ± 0.08 (1.42) | 2.27 ± 0.07 (1.54) | 2.24 ± 0.11 (1.51) | 2.27 ± 0.05 (1.49) |
| Family total income (year 1), $ | 41,853 ± 2451 (29,522) | 53,277 ± 2426 (43,716) | 68,598 ± 6678 (51,653) | 51,845 ± 2078 (39,660) |
| Duration of T2D, years | 13.10 ± 0.45 (11.51) | 13.14 ± 0.54 (10.54) | 13.64 ± 0.86 (11.31) | 13.22 ± 0.31 (11.13) |
| SF-12 MCS | 48.86 ± 0.63 (38.63) | 49.56 ± 0.62 (42.39) | 50.76 ± 1.00 (41.32) | 39.63 ± 0.48 (40.72) |
| SF-12 PCS | 38.46 ± 0.71 (38.63) | 40.17 ± 0.68 (42.39) | 40.76 ± 1.00 (41.32) | 39.63 ± 0.48 (40.72) |
| Health expenditures at year 1 | | | | |
| Total Rx expenditures, $ | 3821.9 ± 317.1 (2222.0) | 4443.3 ± 357.4 (2908.4) | 4996.0 ± 382.2 (3573.8) | 4313.2 ± 228.2 (2897.0) |
| Total OOP Rx cost (all medicines), $ | 661.9 ± 158.5 (262.1) | 839.6 ± 59.6 (526.8) | 1316.4 ± 127.5 (923.1) | 861.7 ± 69.7 (475.4) |
| Total OOP cost, $ | 1296.4 ± 204.4 (507.7) | 1443.6 ± 106.8 (963.4) | 1885.1 ± 139.3 (1466.9) | 1470.6 ± 92.7 (886.3) |
| Total healthcare expense, $ | 11,547.0 ± 938.2 (5968.5) | 12,609.6 ± 1023.9 (6236.9) | 10,737.3 ± 958.6 (7258.1) | 11,862.8 ± 629.7 (6418.9) |
| Number of nights in hospital, n | 0.90 ± 0.16 (0.00) | 1.19 ± 0.25 (0.00) | 0.75 ± 0.19 (0.00) | 1 ± 0.13 (0.00) |
| Adherence to medication | | | | |
| MPR for most costly medication class (%) | 78.4 ± 1.4 (90.9) | 69.1 ± 1.6 (73.7) | 62.5 ± 2.8 (63.8) | 71.3 ± 1.2 (74.6) |
| MPR for least costly medication class (%) | 77.5 ± 1.7 (88.5) | 76.5 ± 1.6 (95.3) | 76.7 ± 2.3 (86.9) | 76.9 ± 1.0 (89.9) |
| ΔMPR (%) | 0.9 ± 1.4 (−0.3) | −7.5 ± 1.9 (0.0) | −14.2 ± 3.7 (−13.7) | −5.6 ± 1.3 (0.0) |
$862 \pm 70$ (median $475$) and OOP cost for all healthcare was $1471 \pm 93$ (median $886$).

**Summary Statistics of OOP Costs and Adherence Behavior**

As described in Table 1, the mean OOP cost per day for the most costly medication was $1.1 \pm 0.2$ per day (median $0.3$) and for the
least costly medication was $0.12 ± $0.01 per day (median $0.09). The mean OOP cost difference (ΔOOP) was $0.99 ± $0.18 per day (median $0.16). Across the entire sample, the mean MPR for the most costly medication was 71.3% ± 1.2% (median 74.6%) and for the least costly medication was 76.9% ± 1.0% (median 89.9%). In the least ΔOOP group, the adherence to the least costly medication was 77.5% and for most costly medication it was 78.4%, with only a mean ΔOOP cost of $0.02/day. As the mean ΔOOP cost increased to $0.41/day for the moderate ΔOOP group, the MPR dropped very slightly to 76.5% for least costly medication but considerably more for the most costly medication to 69.1%. In the high ΔOOP group, with a mean ΔOOP of $4.34/day, the MPR for least costly medication remained approximately the same (76.7%), whereas the MPR for most costly medication dropped to 62.5%. Regardless of the increase in ΔOOP cost, MPR for least costly medication averaged about 77% over a year for all three ΔOOP groups, but, that for the most costly medication dropped by 15.9% from the least ΔOOP group to the highest ΔOOP group. The mean ΔMPR for the least ΔOOP group was near zero (0.9% ± 1.4%) indicating equal adherence to both medications, whereas the corresponding adherence difference for the highest ΔOOP group was – 14.2% ± 3.7%.

The ΔOOP cost is skewed toward $0, with few outliers of very high ΔOOP cost difference (Fig. 1a). In contrast, ΔMPR is relatively symmetric, with 28.5% of patients having the same adherence level to their medications regardless of OOP cost (Fig. 1b). Overall, patients tended to lean toward more adherent to their least costly medication with 42.5% patients more adherent to the least costly medication and 29.0% patients more adherent to the most costly medication. About 34.9% of the patients had ≥ 80% MPR for both least and most costly medications, 28.7% had < 80% MPR for both medications, 22.9% had ≥ 80% MPR for only the least costly medication, and 13.5% had ≥ 80% MPR for only the most costly medication. Therefore, 57.8% of the patients exceeded 80% MPR for their least costly medication and 48.4% had ≥ 80% MPR for the most costly medication.

### Primary Analysis: Correlation Between ΔOOP Costs and ΔMPR

The ΔOOP was significantly negatively correlated with the ΔMPR (Fig. 2). On the basis of the significant slope ($p < 0.001$), for every increase in ΔOOP decile (ΔOOP/day increase ranging from $0.03 to $0.93 increase in each decile), the ΔMPR decreased on average by 1.8% ± 0.4% ($p < 0.001$). Noting the shaded region of Fig. 2 along the x-axis, the mode within each level of ΔOOP, the mode of patients adhered to both medications approximately equally, although this number decreased as the ΔOOP in each decile increased.

As the ΔOOP increased, the variability in ΔMPR markedly increased (Fig. 3), which was confirmed by the post hoc Brown–Forsythe test ($p < 0.001$). Figure 3 illustrates the ΔOOP groups that were defined post hoc. The first three OOP cost deciles were categorized as least ΔOOP group because within this group the boxes that represent the middle 50% of the distribution (25th to 75th percentile) were relatively narrow; most individuals had similar adherence to their most and least costly medication if their OOP cost differed by less than $0.078/day (Fig. 3 inset table). The next five deciles were categorized as moderate ΔOOP group whose ΔOOP costs ranged from $0.08 to $1.08 per day. The median ΔMPR was near zero, with 29.0% of patients having no difference in adherence, 28.5% being more adherent to the most costly medication, while 42.5% of them being more adherent to the least costly medication. However, the sizes of the boxes in the plot increased compared to the least ΔOOP group, indicating increased variability in ΔMPR. Compared to the least ΔOOP group, while there was no distinct shift to the least costly medication, adherence choices became more variable in response to the increase in the ΔOOP between the medications and fewer patients had equal adherence to both medications. The last two cost deciles were categorized as large ΔOOP group which has patients with ΔOOP costs exceeding $1.08/day or $32.98/month. The mean and median ΔMPR in this group decreased, shifting adherence towards the least costly medication away from the most costly medication.
Factors Affecting Adherence Behavior

A multinomial logistic regression identified three variables that significantly affected adherence behavior. These included the ratio of daily OOP between the least and the most costly medications (relative OOP; \( p = 0.002 \)), whether most costly medication type was an insulin and/or fixed-dose combination (FDC) \( (p = 0.006) \), and Short-Form-12 physical component summary (PCS) score in previous year (baseline health status; \( p = 0.006; \) Table 2). The Supplementary Material illustrates that CART of ΔMPR dichotomized the classes into insulin or FDC versus all others, hence their grouping here. In contrast to the ratio, the sum of the OOP costs per day for both the least and most costly medications (OOP sum/total) was not statistically significant overall \( (p = 0.13) \), it played a significant role defining the not-adherent-to-either medication group. Other variables that did not significantly affect adherence when added to the multinomial model included education \( (p = 0.18) \), number of diabetes medications \( (p = 0.84) \), primary cardiovascular (CV) disease risk with hypertension \( (p = 0.24) \), primary CV disease risk with hypercholesterolemia \( (0.87) \), secondary CV risk \( (p = 0.20) \), family income \( (p = 0.94) \), age \( (p = 0.69) \), gender \( (p = 0.34) \), and insurance status \( (p = 0.49) \).

For those patients adherent only to the least costly medication, as the relative OOP increased by 10% \( (i.e., \Delta \text{OOP reduced/cost difference becomes smaller}) \), the odds of only being adherent to the least costly medication decreases \( (\text{odds ratio (OR)} [95\% \text{ CI}], 0.90 [0.84–0.96]; \ p = 0.002) \). For those patients in the less frequent “adherent to only the most costly medication” group, the key factor associated with this behavior is whether the most costly medication is an insulin or FDC, which increased the odds of being in this group to 1.91 \( (\text{OR [95\% CI]} 1.20–3.04; \ p = 0.007) \). The odds of being a patient not adherent to either medication increased as the OOP sum doubled \( (\text{OR [95\% CI]} 1.40 [1.06–1.86]; \ p = 0.02) \). As the previous year’s SF-12 PCS score increased, indicating better patient perception of their baseline physical functional status, the odds of being non-adherent to both medications increased \( (\text{OR [95\% CI]} \ 5\text{-point PCS increase, 1.16 [1.07–1.26]; } \ p < 0.001) \). Replacing this significant variable with a five-level self-reported health score or a comorbidity score in the regression yielded the same result. Although the increase in SF-12 PCS score did not reach statistical significance, it appeared to be associated with improved odds of adherence to the least costly medication \( (\text{OR} = 1.08; \ p = 0.08) \). Further details are provided in the Supporting Material.

DISCUSSION

This study examined the association between the adherence behavior to combination therapy and the relative difference in the OOP costs for patients with T2D on multiple antidiabetic medications. As expected, the association was negative, with adherence shifting towards least costly medication as the OOP cost difference increased. As per our expectation, as the difference in OOP costs increased, we observed a proportional/gradual decrease in adherence to the more costly medication. However, we found further nuances to this association. In combination therapy, adherence was similar for two medications with similar OOP costs. However, as OOP cost difference increased (between $2.39/month and $33/month), even though the median difference adherence did not shift adherence away from the most costly medication, the between-medication variability of adherence increased. At this moderate OOP cost
patients might favor either the most or least costly medication depending on patient characteristics such as medication prescribed and health status. However, beyond $33/month of ΔOOP cost, we observed our a priori expectation of reduced adherence for the most costly medication. High OOP cost difference eventually emerged as the dominating factor leading more patients to significantly reduce consumption of the most costly medications, disregarding the intended complementary nature of the medications in a combination therapy.

The main justification for higher demand-side cost-sharing is that the “moral hazard” (incentive to take unusual risks because others bear the economic consequences) that comes with generous insurance encourages overconsumption of medical care [7, 29, 30]. The widely adopted remedy for this is to trade off the benefits of risk spreading (insurance) against the costs of moral hazard by transferring risk to insured individuals using the form of higher demand-side cost-sharing [6]. This theory further postulated that cost-sharing would only discourage the consumption of the low efficiency or unnecessary medical services because the rational utility-maximizing consumers would demand medical care to the point where marginal costs equals marginal benefits. But this remedy could result in unintended effects if the design of cost-sharing does not recognize the heterogenous disease characteristics, treatment patterns, or patient needs. As T2D is a progressive disease, combination therapy becomes necessary for glycemic control, especially for patients with a longer duration of diabetes [15].

Fig. 2 Correlation between differences in OOP and adherence between patients’ least and most costly medication. The thick blue line indicates the regression line along with the blue shaded part indicating 95% confidence interval. The gray-to-blue shading corresponds to the weighted n on the second y-axis. MPR medication possession ratio, OOP out-of-pocket.
Many patients eventually need to use injectables in combination with oral medications. For them, the medications in the prescribed combination regimen are complements to each other rather than substitutes, i.e., all components in the combination therapy are considered effective and necessary. Therefore, adherence to all the components in a combination regimen should be equally observed for optimal glucose control.

Yet, our study highlights that patients made heterogeneous treatment adherence decisions when faced with different OOP costs for their medications. Nearly a quarter of the studied patients appeared to forgo the more costly medication in favor of the least costly medication when the OOP cost difference was large enough, despite the necessity to adhere to all the medicines of the combination therapy. For these patients, OOP cost consideration appeared to outweigh the medical necessity as their relative OOP cost was the most significant variable that impacted the adherence decisions. One possible explanation could be that some patients did not have the necessary experience or information for their medical decisions, as suggested by Rice [9], and therefore treated the complementary medicines in a combination therapy as if they were perfect substitutes. The RAND Health Insurance Experiment (HIE) [31] also found that patients reduced their consumption of both effective and less effective medical services when faced with cost-sharing. This elastic adherence behavior to OOP costs is concerning and could have negative health consequences for this subgroup of patients if treatment regimen is not adhered to as prescribed.

![Fig. 3 Variability in difference in the adherence between medication class (ΔMPR) by decile of out-of-pocket cost difference (ΔOOP). The box indicates the 25th and 75th percentiles, the blue line within the box indicates the median, and the diamond indicates the mean. MPR medication possession ratio, OOP out-of-pocket](image-url)
It is evident that a third of the studied patients were equally adherent to their medications regardless of the OOP cost difference between medications, indicating the medications were treated as complements as they should be in combination therapies. More than one-eighth of the study population may have behaved like informed consumers who considered the medical effectiveness when making medical consumption decisions, as Rice [9] had theorized. For these patients adherent only to the most costly medication, the sole significant factor associated with this behavior was whether the most costly medication is an insulin or FDC. They might be driven by the perception that certain medication like insulin was important for their health [23]. Nevertheless, most costly medications were not necessarily the most effective and it was not optimal for this group of patients to be non-adherent to their combination therapy as prescribed by physicians.

In contrast to other patients, for nearly a third of the study population who were non-adherent to either of their medications in the combination therapy the relative OOP costs between the medications was not relevant. Instead, the overall OOP burden influenced their choice. Another finding among this group of patients was that patients with better self-perception of physical health had increased odds of being non-adherent to both medications. With perceived better health, patients might not feel the urgency to maintain

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adherence to medications at the 80% level, particularly if the overall medication burden was high. The lack of notable symptoms among patients with chronic conditions such as osteoporosis or diabetes could create a perception of low or even no risk of complications if medications were not used as prescribed [32, 33]. It is therefore important to find ways to improve the adherence among these patients by reducing the overall OOP cost burden and through disease education. Reduced cost-sharing for preventive drugs may benefit low-income patients with diabetes in high-deductible health plans with a health savings account [34].

Our study showed that, in the USA, medication adherence in patients with T2D on combination therapy has not been ideal to begin with, let alone overconsumption of the medications. The adherence to the most costly medication averaged 78.4% within the group with near zero cost difference, while it was 62.5% for the high OOP cost difference group. More concerning is that only one-third of the patients were adherent to both medications in the combination therapy. Therefore, while the high cost-sharing strategies may be necessary to curb resource utilization in some area, the attempt of cost control scheme for prescription medications among patients with T2D (especially among those on combination therapy) might be misplaced. While previous reports clearly suggest that reduced cost-sharing for targeted prescription drugs is significantly associated with improved adherence [35], our study suggests that reducing the relative OOP cost difference between medications would matter significantly for the adherence to combination therapy. Given that the adherence behavior started to vary when difference in OOP costs between medications was as low as $2.39/month, minimizing the OOP cost difference in tiered copay insurance design (which created the OOP cost difference between medications) or moving away from high-deductible design (which caused the high OOP cost difference during the early period in a given insured year) might be needed to improve adherence of patients with T2D to combination therapy.

Our study differed in several aspects from other studies in the research area of adherence and OOP cost relationship. First, each of the patients (observation) in this study was simultaneously faced with two OOP costs and two adherence decisions while in most other studies each patient (observation) was faced with one price and one adherence decision. Secondly, most other studies examined the relationship between adherence and the OOP cost using cross-sectional data, which is unable to account for the unobserved interpatient heterogeneity. Unaccounted variables such as the safety profile of prescribed medicines, general preferences for pharmaceutical, the seriousness of the illness, and the perceived effectiveness of the relevant medication may be important predictors of pharmaceutical demand. Failing to account for important predictors can lead to omitted variable bias [36, 37]. In our study, we reduced both observed and unobserved heterogeneity by using a paired within-patient approach for the primary analysis. To our knowledge, this is the only study that discusses how adherence responds to differences in OOP costs. Thirdly, we employed a design-based approach in selecting the variables that matter to the adherence pattern. Importantly, the variable list produced by this approach is consistent with the prediction of the economic theory [7] that demand for medications is determined by the relative prices (i.e., relative OOP costs) as well as a few other non-cost factors such as perceived health status and proxy for medication effectiveness. The use of relative OOP cost in the multinomial logistic regression reflected the trade-offs that patients had to consider in making treatment decisions. In most other retrospective analysis of drug demand, it was the OOP costs of the selected treatment rather than the relative OOP costs associated with the treatments that were considered [38].

Much evidence suggests that adherence is influenced by age, duration of therapy, education, income, prescription factors, and pill burden [22]. In contrast, we found factors such as number of diabetes medications, family income, age, gender, and insurance status did not significantly influence the adherence behavior. Education was an exception, which
was marginally significant, consistent with other studies. These differences are likely due to the paired within-patient methodological approach taken in this study as well as the characteristics of the patients with T2D on combination therapy.

Limitations

Though patients selected for this study did not change their number of antidiabetic drug classes over the 2-year period (this was designed to minimize medication additions or discontinuations), we could not directly observe whether this was the case. In addition, we followed a published framework that estimated OOP cost derived from the last purchase of the calendar year. As a result of the complexity of plans in which collective OOP costs can change at any point during the year, we believe using the last OOP cost provides a reasonable approach, but we could not determine how often that reflects the reality of their potential subsequent OOP cost. In addition, this study utilized self-reported information, which may be subject to reporting bias. We excluded patients who did not complete 2 years of MEPS participation, so those who died, dropped out, or were institutionalized during participation were not included. We cannot assess the impact of the burden of personal expenses outside the healthcare sector.

CONCLUSIONS

In the context of patients with T2D taking combination therapy, variability in medication adherence increases initially until a cost threshold, beyond which the adherence to more costly medication decreases. This suggests that adherence to combination therapy may be compromised by the heterogeneous sensitivity to the relative OOP costs of medications or to the total burden of OOP costs. There is a need to rethink and recalculate the optimal prescription drug cost-sharing structures for patients with T2D to improve the medication adherence.

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Compliance with Ethics Guidelines. MEPS is a publicly available de-identified data source. MEPS has been reviewed and approved by the Westat IRB, established under a multi-project assurance (MPA M-1531) granted by the Office
for Protection from Research Risks (OPRR). The project is reviewed and the approval renewed annually.

**Data Availability.** Data from the MEPS are freely available from https://meps.ahrq.gov/data_stats/download_data_files.jsp. Codes used for data analysis are available on request from the corresponding author.

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