Research Article

Age at last birth and risk of developing epithelial ovarian cancer: a meta-analysis

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Background: Many epidemiologic studies have explored the association between age at last birth (ALB) and the risk of epithelial ovarian cancer, but the results remain controversial. Methods: A literature search was performed in PubMed, Web of Science, China National Knowledge Infrastructure (CNKI) and WanFang Med Online for relevant articles published up to April 2019. Pooled relative risks (RRs) with 95% confidence intervals (CIs) were calculated using a random-effect model. Dose–response relationship was assessed by restricted cubic spline model. Results: Thirteen articles with 19,959 cases and 2,451,071 participants were included in our meta-analysis, and we found that ALB was negatively associated with epithelial ovarian cancer. The pooled RR (95% CI) of epithelial ovarian cancer for the highest versus the lowest stratification of ALB was 0.77 (0.65–0.91). Furthermore, significantly negative associations were shown in case–control studies (RR: 0.73; 95% CI: 0.60–0.88), studies conducted in North America (RR: 0.71; 95% CI: 0.60–0.84), studies with adjustment for parity (RR: 0.76; 95% CI: 0.63–0.93), studies with adjustment for tubal ligation (RR: 0.74; 95% CI: 0.58–0.94), in the subgroup analysis. In dose–response analysis, the risk of epithelial ovarian cancer decreased nonlinearly with the increase of ALB, and the negative results become significant when ALB was 22.5 years old. Conclusion: This meta-analysis suggested that ALB was negatively associated with the risk of epithelial ovarian cancer. The risk of epithelial ovarian cancer decreased gradually with the ALB for women.

Introduction

Epithelial ovarian cancer is a gynecologic malignancy with fairly high mortality. For most patients with epithelial ovarian cancer, they are diagnosed in advanced stages and have a poor prognosis [1]. According to cell histology, epithelial ovarian cancer is classified into four main subtypes: serous tumor (70–80%), clear cell tumor (5–10%), endometrioid tumor (10%) and mucinous tumor (3–4%) [1,2]. The menstrual, reproductive and hormonal factors are closely associated with the risk of developing epithelial ovarian cancer [3–8], and the role of some of these factors are now well investigated. Factors connected with a reduced risk of epithelial ovarian cancer include the use of oral contraceptive (OC), high parity, breastfeeding and a history of tubal ligation [7,9–12]. In the other hand, endometriosis, the use of menopausal hormone and later menopause are risk factors for epithelial ovarian cancer [13–15].

Age at last birth (ALB) plays a vital role in the pathogenesis of gynecological cancer, which maybe because of the prenatal hormone changes [16]. Study has found that a later ALB is associated with a lower risk of endometrial cancer [17]. However, the association between ALB and the risk of epithelial ovarian cancer is still controversial. Numerous epidemiological studies have been done to explore the association between ALB and the risk of epithelial ovarian cancer [3,18–29]. Among these studies, eight studies [18–20,22–24,26,29] showed no apparent association between ALB and epithelial ovarian cancer; four [21,25,27,28] showed a negative association; and one [3] showed a positive association. Thus, we conducted this meta-analysis to explore the association between ALB and risk of epithelial ovarian cancer systematically.

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Materials and methods
Preferred Reporting Items for Systematic reviews and Meta-Analyses (PRISMA) guidelines were consulted in this analysis [30].

Literature search strategy
All relevant studies (up to April 2019) were identified by a comprehensive search in PubMed, Web of Science, Chinese National Knowledge Infrastructure (CNKI) and WanFang Med Online. Search terms included ‘age at last birth’ (or ‘age at last delivery’ or ‘age at last pregnancy’ or ‘age at last child’ or ‘age at last childbirth’ or ‘age at last full-term pregnancy’) and ‘ovarian cancer’ (or ‘oophoroma’ or ‘carcinoma of the ovary’ or ‘ovarian neoplasm’ or ‘ovarian tumor’). We also manually searched the relevant references within included studies to find eligible articles. The detailed steps of the literature selection are shown in Figure 1.

Inclusion criteria
Two investigators (Yanjun Wu and Wenjun Sun) checked carefully all identified articles. If the article met the following criteria, it would be included in our meta-analysis. (1) a case–control or cohort study was published as an original study; (2) the exposure of interest was ALB; (3) the outcome of interest was epithelial ovarian cancer; (4) There was reported effect size (relative risk (RR) or odds ratio (OR) or hazard ratio (HR) or incidence rate ratio (IRR)) and 95% confidence interval (CI) for the association between ALB and epithelial ovarian cancer; (5) We selected the most recent study, if data from the same population were used in multiple articles.

Data extraction
From each eligible article, we extracted the first author’s name, country where the study was conducted, publication year, the type of study design, the follow-up duration of cohort study, age range or mean age at baseline. We also extracted the number of cases and participants, the information about ALB, RRs (we presented all results as RR for simplicity) with their 95% CIs for each stratification of ALB, as well as adjusted factors.

Statistical analysis
The lowest stratification of ALB was considered as the reference category in the majority of studies. When the reference category was not the lowest stratification of ALB [21,23,26], the method proposed by Hamling et al. [31] was used to convert risk estimates. In article [29] that only reported ‘floating absolute risk’ CIs, conventional CIs should be calculated by using the approach of Easton et al. [32]. For articles [27,29] that only reported on the association between ALB and epithelial ovarian cancer split by subgroup (e.g. menopausal status or tumor subtype), a random-effect model (REM) was first used to generate an overall result. Then a pooled measure was calculated as the inverse variance-weighted mean of the logarithm of RR with corresponding 95% CI to assess the strength of association between ALB and the risk of epithelial ovarian cancer. The $I^2$ of Higgins and Thompson was adopted to assess the heterogeneity among studies, and $I^2$ values of 0, 25, 50 and 75% represent no, low, moderate and high heterogeneity [33]. If $I^2 < 50\%$, the fixed-effect model (FEM) would be used to combine study-specific RRs (95% CIs), otherwise, the REM which considers both within-study and between-study variation would be used ($I^2 \geq 50\%$) [34,35]. Meta-regression was used to explore the important potential covariates that might exert substantial impacts on between-study heterogeneity [36]. Subgroup analyses were performed by continent, the type of study design, whether or not the results were adjusted for parity, the use of OC and tubal ligation. Leave-One-Out sensitivity analysis was performed to evaluate the pivotal studies that have substantial impacts on between-study heterogeneity [37]. The influence analysis was performed to assess whether the result could have been affected obviously by a single study. Publication bias was evaluated using the Egger’s test and funnel plot [38].

For dose–response analysis, a two-stage random-effects dose–response meta-analysis [39] was performed. In the first stage, a restricted cubic spline model with three knots at the 25th, 50th and 75th centiles [40] of the levels of ALB was estimated using generalized least square regression, taking into account the correlation within each set of published RRs [41]. Then the study-specific estimates were combined using the restricted maximum likelihood method in a multivariate random-effects meta-analysis [42]. A $P$-value for nonlinearity was calculated by testing the null hypothesis that the coefficient of the second spline is equal to 0.

All statistical analyses were performed with Stata version 15.0 (Stata Corp., College Station, TX, U.S.A.). All reported probabilities ($P$-values) were two-sided, and $P$-values less than 0.05 were considered statistically significant.
Articles identified by searching PubMed (n=168), Web of Science (n=251), CNKI (n=0), and Wan Fang Med Online (n=10)

Additional articles identified by manually searching reference list of included articles (n=4)

334 articles were identified

286 articles were excluded after review of title and abstract

48 articles were read in full-text

35 articles were excluded because:
1: the outcome was ovarian cancer (all types), not epithelial ovarian cancer
2: with duplicate data
4: the outcome was ovarian cancer survival
16: did not evaluate the association of the interest
12: did not report RRs or 95% CIs concerning the association between ALB and epithelial ovarian cancer

13 articles were included in the meta-analysis

Figure 1. Flow chart of the selection of studies included in the meta-analysis
Results

Literature search and characteristic of included studies
Using our search strategy mentioned in ‘material and method’, we identified 334 articles in total. There were 168 articles from PubMed, 251 articles from Web of Science, 10 articles from Wan Fang Med Online and 4 articles from reference list of included articles. We excluded 286 articles after reviewing the title and abstract, and we further excluded 35 articles in the step of full-text article reviewing. Among these 35 articles, the outcome of 4 articles [43–46] was ovarian cancer (not restricted to epithelial ovarian cancer), the outcome of one article was ovarian cancer survival [47], two articles [48,49] had the same population, 16 articles failed to evaluate the association between ALB and epithelial ovarian cancer and 12 articles did not report RRs and 95% CIs for the association of interest. Ultimately, thirteen [3,18–29] articles were included in our meta-analysis. The detailed steps of the article selection were shown in Figure 1. In these included articles, seven studies [19–24,27] were conducted in North America, five [3,18,26,28,29] in Europe and one [25] in Oceania. With regard to the type of study design, ten articles [3,19–22,24–28] were case–control studies and three [18,23,29] were cohort studies. The characteristics of the included studies are shown in Table 1.

Overall association between ALB and epithelial ovarian cancer
Thirteen articles [3,18–29] involving 19,959 cases and 2,451,071 participants were included to evaluate the association between ALB and epithelial ovarian cancer. Among these studies, eight studies [18–20,22–24,26,29] showed no apparent association between ALB and epithelial ovarian cancer; four [21,25,27,28] showed a negative association; and one [3] showed a positive association. The pooled RR of epithelial ovarian cancer for the highest versus the lowest stratification of ALB was 0.77 (95% CI: 0.65–0.91; $I^2 = 73.90\%$, $P_{\text{for heterogeneity}} < 0.001$). The detailed results are shown in Figure 2.

In dose–response analysis, data from nine studies [3,20–23,25–28] were used and a nonlinear association was found ($P_{\text{nonlinearity}} = 0.004$) between ALB and risk of epithelial ovarian cancer. The risk of epithelial ovarian cancer decreased gradually with the ALB, and the results become significant when ALB was 22.5 years old. The RRs (95% CIs) of epithelial ovarian cancer risk for the 25th, 50th and 75th centiles of the levels of ALB were 0.99 (0.99–0.99), 0.95 (0.94–0.96) and 0.88 (0.86–0.90) for 22.5, 27.5 and 32.5 years old, respectively. The detail results are shown in Figure 3.

Data from five studies [22,24,26,28,29] were used to explore the association between ALB and the histologic subtypes of epithelial ovarian tumor. The pooled RR (95%CI) for the highest versus the lowest stratification of ALB was 0.80 (0.62–1.03) for serous tumors, 0.91 (0.77–1.07) for mucinous tumors, 0.72 (0.42–1.24) for endometrioid tumors and 0.83 (0.63–1.09) for clear cell tumors. The characteristics of the included studies and the detailed results are shown in Table 2.

Subgroup analysis
In the subgroup analysis by continent where the study was conducted, the pooled RRs (95% CIs) were 0.71 (0.60–0.84) and 0.86 (0.65–1.14) for the studies conducted in North America and Europe, respectively. In the subgroup analysis by study design, the pooled RRs (95% CIs) for case–control and cohort studies were 0.73 (0.60–0.88) and 0.92 (0.85–1.00), separately. According to whether the results were adjusted for parity or not, the pooled RRs (95% CIs) were 0.76 (0.63–0.93) and 0.82 (0.61–1.11) for studies with adjustment and without, respectively. According to whether the results were adjusted for the use of OC or not, the pooled RRs (95% CIs) for studies with adjustment and without were 0.81 (0.63–1.03) and 0.66 (0.59–0.74), severally. In the subgroup analysis of whether the results were adjusted for the tubal ligation or not, the pooled RRs (95% CIs) for studies with adjustment and without were 0.74 (0.58–0.94) and 0.81 (0.62–1.04), respectively. Detailed results of subgroup analysis are shown in Table 3.

Meta-regression
The results of multivariable meta-regression showed that publication year ($P=0.401$), continent where the study was conducted ($P=0.900$), the type of study design ($P=0.280$), whether or not the results were adjusted for parity ($P=0.831$), the use of OC ($P=0.580$) and tubal ligation ($P=0.655$) had no significant impact on the heterogeneity.

Sensitive analysis
Leave-One-Out sensitivity analysis was performed to evaluate the key studies that had substantial impacts on between-study heterogeneity. Two studies [3,29] were found to contribute to between-study heterogeneity. After excluding the two studies, the heterogeneity decreased to 0.0% ($P_{\text{for heterogeneity}} = 0.680$) and the result was still significant 0.67 (95% CI 0.61–0.73).
Table 1 Characteristics of the included studies for ALB with risk of the epithelial ovarian cancer

| First author       | Country (year) | Age range or Mean age (Case/control) | Study design | Exposure (ALB) | Case Participants | RR(95% CI) | Adjustment for covariates                                                                 |
|-------------------|----------------|--------------------------------------|--------------|----------------|-------------------|------------|------------------------------------------------------------------------------------------|
| Gaitskell         | U.K. (2018)    | 56.1                                 | Cohort (14.6)| <25            | 2043              | 0.96 (0.89–1.04)| Age, region, tubal ligation, hysterectomy, family history of breast cancer, the use of OC, use of menopausal hormones, BMI, smoking, SES, parity. |
|                   |                | 25–29                                |              |                |                   |            |                                                                                           |
|                   |                | ≥30                                  |              |                | 1823              | 0.93 (0.85–1.01)|                                                                                           |
|                   |                | 25–29                                |              |                | 3454              | 0.85 (0.80–0.91)|                                                                                           |
| Sköld             | Nordic (2018)  | 19–85                                | Case–control | <25            | 1455              | 118821     | 0.76 (0.71–0.82)                                                                                   |
|                   |                | 25–29                                |              |                |                   |            |                                                                                           |
|                   |                | ≥30                                  |              |                | 5556              | 0.64 (0.56–0.72)|                                                                                           |
| Moorman           | America (2008) | 20–74                                | Case–control | <25            | 190               | 1543       | 0.76 (0.71–0.82)                                                                                   |
|                   |                | 25–29                                |              |                |                   |            |                                                                                           |
|                   |                | ≥30                                  |              |                | 1823              | 0.93 (0.85–1.01)|                                                                                           |
| Soegaard          | Denmark (2007) | 35–79                                | Case–control | <25            | 100               | 1927       | 0.66 (0.48–0.91)                                                                                   |
|                   |                | 25–29                                |              |                |                   |            |                                                                                           |
|                   |                | ≥30                                  |              |                | 222               | 0.68 (0.45–1.01)|                                                                                           |
| Whitman           | Australia (2003) | 18–79                              | Case–control | <25            | 117               | 1343       | 0.76 (0.52–1.10)                                                                                   |
|                   |                | 25–29                                |              |                |                   |            |                                                                                           |
|                   |                | ≥30                                  |              |                | 79                | 0.54 (0.38–0.77)|                                                                                           |
| Tung              | America (2003) | 18+                                  | Case–control | <25            | NA                | NA         | 0.76 (0.52–1.10)                                                                                   |
|                   |                | 25–29                                |              |                |                   |            |                                                                                           |
|                   |                | ≥30                                  |              |                | 171               | 0.56 (0.37–0.84)|                                                                                           |
| Vachon            | America (2002) | 56–81                                | Cohort (13)  | ≤29            | 66                | 31377      | 0.90 (0.60–1.30)                                                                                   |
|                   |                | 28–30                                |              |                |                   |            |                                                                                           |
|                   |                | ≥34                                  |              |                | 90                | 0.57 (0.26–0.90)|                                                                                           |
| Titus-Ernstoff    | America (2001) | 20–74                                | Case–control | <25            | 69                | 796        | 0.76 (0.51–1.13)                                                                                   |
|                   |                | 25–29                                |              |                |                   |            |                                                                                           |
|                   |                | ≥35                                  |              |                | 119               | 0.62 (0.40–0.99)|                                                                                           |
| Cooper            | America (1999) | 18–79                                | Case–control | <25            | 199               | 4060       | 0.87 (0.60–1.20)                                                                                   |
|                   |                | 26–30                                |              |                |                   |            |                                                                                           |
|                   |                | ≥31                                  |              |                | 201               | 0.64 (0.43–0.96)|                                                                                           |
| Salazar-Martinez  | Mexico (1999)  | 52.8/54.6                            | Case–control | ≤25            | 36                | 752        | 0.71 (0.53–0.96)                                                                                   |
|                   |                | 26–35                                |              |                |                   |            |                                                                                           |
|                   |                | ≥36                                  |              |                | 28                | 0.65 (0.43–1.10)|                                                                                           |
| Godard            | Canada (1998)  | 20–84                                | Case–control | 17–29          | NA                | NA         | 0.87 (0.40–1.80)                                                                                   |
|                   |                | 26–35                                |              |                |                   |            |                                                                                           |
|                   |                | ≥36                                  |              |                | 20                | 0.87 (0.40–1.80)|                                                                                           |

Continued over
Table 1 Characteristics of the included studies for ALB with risk of the epithelial ovarian cancer (Continued)

| First author | Country (year) | Age range or Mean age (Case/control) | Study design (years of follow up) | Exposure (ALB) | Case | Participants | RR(95% CI) | Adjustment for covariates |
|--------------|---------------|--------------------------------------|----------------------------------|----------------|------|--------------|------------|---------------------------|
| Albrektsen   | Norway (1996) | 20–56                                | Cohort (16.4)                    | 30–44          | <25  | 218          | 1145076    | 0.63 (0.34–1.15) 1 Age, birth-cohort, parity, age at first birth, time since last birth. |
|              |               |                                      |                                  | 25–29          | 416  | 1            |            | 0.93 (0.77–1.12) |
|              |               |                                      |                                  | 30–34          | 256  | 1            |            | 0.84 (0.66–1.08) |
|              |               |                                      |                                  | ≥35            | 81   | 1            |            | 0.77 (0.54–1.11) |
| Tavani       | Italy (1993)  | 15–44                                | Case–control                     | <25            | 23   | 815          | 1          | 1.30 (0.70–2.40) Age, education, family history, parity, number of abortions, the use of OC. |
|              |               |                                      |                                  | 25–29          | 48   | 1            |            | 2.40 (1.30–4.50) |
|              |               |                                      |                                  | ≥30            | 52   | 1            |            | 2.40 (1.30–4.50) |

Abbreviations: BMI, body mass index; NA, not available; SES, socioeconomic status.

Figure 2. Forest plot of ALB and the risk of epithelial ovarian cancer
The size of gray box is positively proportional to the weight assigned to each study, and horizontal lines represent the 95% CIs. The RR (95% CI) in every article is the RR (95% CI) of epithelial ovarian cancer for the highest versus the lowest stratification of ALB.

Influence analysis and publication bias
The result of influence analysis showed that no study had excessive influence on the pooled RR of epithelial ovarian cancer for the highest versus the lowest stratification of ALB. No evidence of significant publication bias was found by the visual inspection of the funnel plot (Figure 4) and Egger’s test (P=0.581).
Figure 3. The dose–response analysis between ALB and the risk of epithelial ovarian cancer with restricted cubic splines in a multivariate random-effects dose–response model
The solid line and the long dash line represent the estimated RRs and its 95% CIs. Short dash line represents the linear relationship.

Table 2 RRs and 95% CIs for the association between ALB and risk of the major epithelial ovarian cancer subtypes

| Author (year)   | Exposure (ALB) | Serous | Mucinous | Endometrioid | Clear cell |
|-----------------|----------------|--------|----------|--------------|------------|
| Gaitskell (2018)| ≥30/<25        | 0.94 (0.85–1.04) | 0.91 (0.74–1.12) | 0.96 (0.74–1.24) | 0.78 (0.56–1.08) |
| Titus-Ernstoff (2001)| ≥30/<30 | Borderline 1.10 (0.60–2.00) | 0.80 (0.40–1.60) | 1.00 (0.70–1.50) |
| Soegaard (2007) | ≥30/<25        | 0.64 (0.39–1.04) | 0.74 (0.25–2.18) | 0.71 (0.27–1.89) |
| Sköld (2018)    | ≥40/<25        | 0.58 (0.48–0.70) | 0.93 (0.65–1.34) | 0.39 (0.24–0.65) | 1.01 (0.57–1.79) |
| Tung (2003)     | >34/<28        | Borderline 1.10 (0.40–2.70) | 0.80 (0.30–2.00) | |
| Pooled RR (95% CI) | The highest versus the lowest stratification | 0.80 (0.62–1.03) | 0.91 (0.77–1.07) | 0.72 (0.42–1.24) | 0.83 (0.63–1.09) |

Discussion

This meta-analysis assessed the association between ALB and epithelial ovarian cancer. Findings from our study provided evidence of a negative association between ALB and the risk of epithelial ovarian cancer. In the subgroup analysis, the significantly negative associations between ALB and epithelial ovarian cancer were observed in studies conducted in North America, studies with adjustment for parity, studies with adjustment for tubal ligation and case-control studies, respectively. In dose–response analysis, the risk of epithelial ovarian cancer decreased nonlinearly with the increase of ALB. We did not find a statistically significant association between ALB and epithelial ovarian cancer by different histological types, although the point estimates suggested a reduction in risk. This may be in part due to the limited number of studies with data for tumor subtypes.

The protective effect of ALB on epithelial ovarian cancer is closely related to the changes during pregnancy, and the changes include the cessation of ovulation cycle, the decrease of gonadotropin and estrogen level as well as the increase of progesterone level. During the ovulatory cycle, the repeated ruptures of the ovarian epithelium accompanied...
Table 3 Summary of subgroup results for association of ALB with risk of the epithelial ovarian cancer

| Stratification                                      | Number of studies | RR (95% CI)      | I^2, % | P for heterogeneity |
|-----------------------------------------------------|-------------------|------------------|-------|---------------------|
| All studies                                         | 13                | 0.77 (0.65–0.91) | 73.90%| <0.001              |
| Continent where the study was conducted             |                   |                  |       |                     |
| North America                                       | 7                 | 0.71 (0.60–0.84) | 0.00% | 0.488               |
| Oceania                                             | 1                 | 0.57 (0.36–0.90) |       |                     |
| Europe                                              | 5                 | 0.86 (0.65–1.14) | 88.70%| <0.001              |
| The type of study design                            |                   |                  |       |                     |
| Case–control study                                  | 10                | 0.73 (0.60–0.88) | 58.30%| 0.010               |
| Cohort study                                        | 3                 | 0.92 (0.85–1.00) | 0.00% | 0.598               |
| Whether the results were adjusted for parity or not |                   |                  |       |                     |
| Yes                                                 | 10                | 0.76 (0.63–0.93) | 80.00%| <0.001              |
| No                                                  | 3                 | 0.82 (0.61–1.11) | 0.00% | 0.617               |
| Whether the results were adjusted for the use of OC or not |                   |                  |       |                     |
| Yes                                                 | 7                 | 0.81 (0.63–1.03) | 76.90%| <0.001              |
| No                                                  | 6                 | 0.66 (0.59–0.74) | 0.00% | 0.675               |
| Whether the results were adjusted for tubal ligation or not |                   |                  |       |                     |
| Yes                                                 | 5                 | 0.74 (0.58–0.94) | 72.20%| 0.006               |
| No                                                  | 8                 | 0.81 (0.62–1.04) | 63.30%| 0.008               |

RR (95% CI) represent the relative risk (95% CI) of epithelial ovarian cancer for the highest versus the lowest stratification of ALB.

Figure 4. Funnel plot with pseudo 95% confidence limits for the analysis of ALB and risk of epithelial ovarian cancer

The RR (95% CI) in every article is the RR (95% CI) of epithelial ovarian cancer for the highest versus the lowest stratification of ALB.

by estrogen-rich environmental exposure and rapid proliferative repair suggested by ‘Incessant ovulation hypothesis’ and the increase of inflammatory response suggested by ‘Inflammation hypothesis’ can increase the possibility of ovarian epithelial cell malignant transformation [50,51]. The ‘Gonadotropin hypothesis’ hold that the excessive gonadotropin and estrogen secretion may contribute to the deterioration of ovarian epithelial cells [52]. Thus, any
factors (such as pregnancy) that reduce the number of ovulation cycle or the level of gonadotropin can lower the risk of epithelial ovarian cancer. However, these hypotheses are insufficient to explain the phenomenon that the later a woman has her last child, the lower her ovarian cancer risk is. A less well-recognized theory, the ‘ovarian clearance hypothesis’ suggests that the precancerous cells can be cleared by apoptosis under the action of elevated progesterone level during pregnancy [21,53–55]. On average, women with older ALB have more deteriorating ovarian epithelial cells than women with younger ALB. Thus the high levels of hormones during pregnancy may provide a greater benefit for women who have their last birth at an older age, in terms of reducing the risk of epithelial ovarian cancer [25,56].

Between-study heterogeneity is common in meta-analyses [57], and it is essential to explore the sources of between-study heterogeneity. In this meta-analysis, high between-study heterogeneity with $\chi^2 = 73.90\%$ was found. Thus, we used the REM that considered both within-study and between-study variation to calculate the pooled RRs with 95% CIs. Beyond that, we also conducted a Leave-One-Out sensitivity analysis to reduce the heterogeneity. Two studies [3,29] were found to contribute to this high between-study heterogeneity. After excluding them, the heterogeneity reduced to 0.0% and the result remains significant, which suggested that our results were stable and credible. But the meta-regression with the covariates of study design, publication year, continent where the study was conducted, whether or not the results were adjusted for parity, the use of OC and tubal ligation failed to explain the source of heterogeneity. It indicated that some other unknown confounding factors existed in the studies.

There were many advantages in our study. First, most of the studies included in this meta-analysis had adjusted for the potential confounders that had a greater impact on the association between ALB and epithelial ovarian cancer (such as parity and tubal ligation). Second, the reverse association between ALB and epithelial ovarian cancer had no substantial change in the sensitivity analysis, which might suggest that the results were reasonable and convincing. Finally, we further conducted a dose–response analysis to explore the association between ALB and epithelial ovarian cancer.

However, several potential limitations in our study should also be considered. First, some studies including participants who were pre-menopausal might have a risk of reverse causality. Women who had epithelial ovarian cancer at a younger age might have their ovaries removed, and so would not be able to have children at an older age. This might confound the protective effect of a later ALB on epithelial ovarian cancer. Second, the age range at baseline of participant varied in every study. Finally, the adjustment factors were different among studies, and residual confounding could not be eliminated thoroughly.

Conclusion
In summary, results from this meta-analysis indicated that ALB was negatively associated with epithelial ovarian cancer. The results mainly came from case–control studies, and thus more cohort studies were needed to confirm the conclusion.

Competing Interests
The authors declare that there are no competing interests associated with the manuscript.

Author Contribution
Y.W.: designed the study; conducted the statistical analysis; edited the manuscript. W.S.: performed the literature research; extracted the data from included studies. X.X. and W.W.: prepared and reviewed the manuscript. D.Z.: guarantor of integrity of the entire study. All authors approved the final manuscript.

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Abbreviations
ALB, age at last birth; CI, confidence interval; OC, oral contraceptive; REM, random-effect model; RR, relative risk.

References
1 Torre, L.A., Trabert, B. and DeSantis, C.E. (2018) Ovarian cancer statistics, 2018. CA Cancer J. Clin. 68, 284–296, https://doi.org/10.3322/caac.21456
2 Prat, J. (2015) Pathology of cancers of the female genital tract. Int. J. Gynaecol. Obstet. 131, S132–S145, https://doi.org/10.1016/j.ijgo.2015.06.010
3 Tavani, A., Negri, E., Franceschi, S., Parazzini, F. and La Vecchia, C. (1993) Risk factors for epithelial ovarian cancer in women under age 45. Eur. J. Cancer 29A, 1297–1301, https://doi.org/10.1016/0959-8049(93)90077-S

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4 Parazzini, F., Franceschi, S., La Vecchia, C. and Fasoli, M. (1991) The epidemiology of ovarian cancer. Gynecol. Oncol. 43, 9–23, https://doi.org/10.1016/0090-8258(91)90003-N

5 Chialfariello, F., Pelucchi, C., Parazzini, F., Negri, E., Franceschi, S., Talamini, R. et al. (2001) Reproductive and hormonal factors and ovarian cancer. Ann. Oncol. 12, 337–341, https://doi.org/10.1093/annonc/12.5.337

6 Negri, E., Franceschi, S., Tzonou, A., Booth, M., La Vecchia, C., Parazzini, F. et al. (1991) Pooled analysis of 3 European case-control studies: I. Reproductive factors and risk of epithelial ovarian cancer. Int. J. Cancer 49, 50–56

7 Sung, H.K., Ma, S.H., Choi, J.Y., Huang, Y., Ahn, C., Kim, B.G. et al. (2016) The effect of breastfeeding duration and parity on the risk of epithelial ovarian cancer: a systematic review and meta-analysis. J. Prev. Med. Public Health 49, 349–366, https://doi.org/10.3961/jpoph.16.066

8 Wentzensen, N., Poole, E.M., Trabert, B., White, E., Arslan, A.A., Patel, A.V. et al. (2016) Ovarian cancer risk factors by histologic subtype: an analysis from the ovarian cancer cohort consortium. J. Clin. Oncol. 34, 2888–2898, https://doi.org/10.1200/JCO.2016.66.8178

9 Luan, N.N., Wu, Q.J., Gong, T.T., Vogtmann, E., Wang, Y.L. and Lin, B. (2013) Breastfeeding and ovarian cancer risk: a meta-analysis of epidemiologic studies. Am. J. Clin. Nutr. 98, 1020–1031, https://doi.org/10.3945/ajcn.113.11062794

10 Li, D.P., Du, C., Zhang, Z.M., Li, G.X., Yu, Z.F., Wang, X. et al. (2013) Breastfeeding and ovarian cancer risk: a systematic review and meta-analysis of 40 epidemiological studies. Asian Pac. J. Clin. Oncol. 15, 4829–4837, https://doi.org/10.7314/APJC.2014.15.12.4829

11 Bosetti, C., Negri, E., Trichopoulos, D., Franceschi, S., Beral, V., Tzonou, A. et al. (2002) Long-term effects of oral contraceptives on ovarian cancer risk. Int. J. Cancer 102, 262–265

12 Wang, C., Liang, Z., Liu, X., Zhang, Q. and Li, S. (2016) The association between endometriosis, tubal ligation, hysterectomy and epithelial ovarian cancer: results from the Danish MALOVA study. Cancer Epidemiol. Biomarkers Prev. 25, 397–402, https://doi.org/10.1158/1055-9965.EPI-16-0089

13 Pearce, C.L., Templeman, C., Rossing, M.A., Lee, A., Near, A.M., Webb, P.M. et al. (2012) Association between endometriosis and risk of histological subtypes of ovarian cancer: a pooled analysis of case-control studies. Lancet Oncol. 13, 385–394, https://doi.org/10.1016/S1470-2045(11)70404-1

14 Franceschi, S., La Vecchia, C., Booth, M., Tzonou, A., Negri, E., Parazzini, F. et al. (1991) Pooled analysis of 3 European case-control studies of ovarian cancer: II. Age at menarche and at menopause. Int. J. Cancer 49, 57–60, https://doi.org/10.1002/ijc.2910490111

15 Beral, V., Gaitkell, K., Hermon, C., Moser, K., Reeves, G. and Petro, R. (2015) Menopausal hormone use and ovarian cancer risk: individual participant meta-analysis of 52 epidemiological studies. Lancet 385, 1835–1842, https://doi.org/10.1016/S0140-6736(14)61687-1

16 Karim, R., Dang, H., Henderson, W.V., Hodis, H.N., St John, J., Brinton, R.D. et al. (2016) Effect of reproductive history and exogenous hormone use on cognitive function in mid- and late life. J. Am. Geriatr. Soc. 64, 2448–2456, https://doi.org/10.1111/jgs.14658

17 Setiawan, V.W., Pike, M.C., Karageorgi, S., Deming, S.L., Anderson, K., Bernstein, L. et al. (2012) Age at last birth in relation to risk of endometrial cancer: pooled analysis in the epidemiology of endometrial cancer consortium. Am. J. Epidemiol. 176, 269–278, https://doi.org/10.1093/aje/kws129

18 Albrechtsen, G., Heuch, I. and Kvale, G. (1996) Reproductive factors and incidence of epithelial ovarian cancer: a Norwegian prospective study. Cancer Causes Control 7, 421–427

19 Godard, B., Foukles, W.D., Provenccher, D., Brunet, J.S., Tonin, P.N., Mes-Masson, A.M. et al. (1998) Risk factors for familial and sporadic ovarian cancer among French Canadians: a case-control study. Am. J. Obstet. Gynecol. 179, 403–410, https://doi.org/10.1016/S0002-9378(98)70372-2

20 Salazar-Martinez, E., Lazacone-Ponce, E.C., Gonzalez Lira-Lira, G., Escudero-De los Ríos, P., Salmeron-Castro, J. and Hernandez-Avila, M. (1999) Reproductive factors of ovarian and endometrial cancer risk in a high fertility population in Mexico. Cancer Res. 59, 3658–3662

21 Cooper, G.S., Schildkraut, J.M., Whittemore, A.S. and Marchbanks, P.A. (1999) Pregnancy recency and risk of ovarian cancer. Cancer Causes Control 10, 397–402

22 Titus-Ernstoff, L., Perez, K., Cramer, D.W., Harlow, B.L., Baron, J.A. and Greenberg, E.R. (2001) Menstrual and reproductive factors in relation to ovarian cancer risk. Br. J. Cancer 84, 714–721, https://doi.org/10.1054/bjoc.2000.1596

23 Vachon, C.M., Minik, P.J., Janney, C.A., Sellers, T.A., Cerhan, J.R., Hartmann, L. et al. (2002) Association of parity and ovarian cancer risk by family history of breast or ovarian cancer in a population-based study of postmenopausal women. Epidemiology 13, 66–71, https://doi.org/10.1097/00001648-200201000-00011

24 Tung, K.H., Goodman, M.T., Wu, A.H., McDuffie, K., Wilkens, L.R., Kolonel, L.N. et al. (2003) Reproductive factors and epithelial ovarian cancer risk by histologic type: a multietnic case-control study. Am. J. Epidemiol. 158, 629–638, https://doi.org/10.1093/aje/kwe177

25 Whitteman, D.C., Siskind, V., Purdie, D.M. and Green, A.C. (2003) Timing of pregnancy and the risk of epithelial ovarian cancer. Cancer Epidemiol. Biomarkers Prev. 12, 42–46

26 Soegaard, M., Jensen, A., Hodgall, E., Christensen, L., Hodgall, C., Blaakaer, J. et al. (2007) Different risk factor profiles for mucinous and nonmucinous ovarian cancer: results from the Danish MALOVA study. Cancer Epidemiol. Biomarkers Prev. 16, 1160–1166, https://doi.org/10.1158/1055-9137.EPI-07-0889

27 Moorman, P.G., Calingaert, B., Palmieri, R.T., Iversen, E.S., Bentley, R.C., Halabi, S. et al. (2008) Hormonal risk factors for cancer in premenopausal and postmenopausal women. Am. J. Epidemiol. 167, 1059–1069, https://doi.org/10.1093/aje/kwn006

28 Skold, C., Bjorge, T., Ekholm, A., Engeland, A., Gissler, M., Grotmol, T. et al. (2018) Preterm delivery is associated with an increased risk of epithelial ovarian cancer among parous women. Int. J. Cancer 143, 1858–1867, https://doi.org/10.1002/ijc.31581

29 Gaitkell, K., Green, J., Pirie, K., Barnes, I., Hermon, C., Reeves, G.K. et al. (2018) Histological subtypes of ovarian cancer associated with parity and histologic progression in the prospective Million Women Study. Int. J. Cancer 142, 281–289, https://doi.org/10.1002/ijc.31063

30 Moher, D., Liberati, A., Tetzlaff, J. and Altman, D.G. (2010) Preferred reporting items for systematic reviews and meta-analyses: the PRISMA statement. Int. J. Surg. 8, 336–341, https://doi.org/10.1016/j.ijsu.2010.02.007

31 Hamling, J., Lee, P., Weilnort, R. and Ambuhl, M. (2008) Facilitating meta-analyses by deriving relative effect and precision estimates for alternative comparisons from a set of estimates presented by exposure level or disease category. Stat. Med. 27, 954–970, https://doi.org/10.1002/sim.3013

32 Easton, D.F., Peto, J. and Babiker, A.G. (1991) Floating absolute risk: an alternative to relative risk in survival and case-control analysis avoiding an arbitrary reference group. Stat. Med. 10, 1025–1035, https://doi.org/10.1002/sim.4780100703
33 Higgins, J.P. and Thompson, S.G. (2002) Quantifying heterogeneity in a meta-analysis. Stat. Med. 21, 1539–1558, https://doi.org/10.1002/sim.1186
34 Higgins, J.P., Thompson, S.G., Deeks, J.J. and Altman, D.G. (2003) Measuring inconsistency in meta-analyses. BMJ 327, 557–560, https://doi.org/10.1136/bmj.327.7414.557
35 Wu, W., Kang, S. and Zhang, D. (2013) Association of vitamin B6, vitamin B12 and methionine with risk of breast cancer: a dose-response meta-analysis. Br. J. Cancer 109, 1926–1944, https://doi.org/10.1038/bjc.2013.438
36 Higgins, J.P. and Thompson, S.G. (2004) Controlling the risk of spurious findings from meta-regression. Stat. Med. 23, 1663–1682, https://doi.org/10.1002/sim.1752
37 Patsopoulos, N.A., Evangelou, E. and Ioannidis, J.P. (2008) Sensitivity of between-study heterogeneity in meta-analysis: proposed metrics and empirical evaluation. Int. J. Epidemiol. 37, 1148–1157, https://doi.org/10.1093/ije/dyn065
38 Egger, M., Davey Smith, G., Schneider, M. and Minder, C. (1997) Bias in meta-analysis detected by a simple, graphical test. BMJ 315, 629–634, https://doi.org/10.1136/bmj.315.7109.629
39 Orsini, N., Li, R., Wolk, A., Khudyakov, P. and Spiegelman, D. (2012) Meta-analysis for linear and nonlinear dose-response relations: examples, an evaluation of approximations, and software. Am. J. Epidemiol. 175, 66–73, https://doi.org/10.1093/aje/kwr265
40 Harrell, Jr, F.E., Lee, K.L. and Pollock, B.G. (1988) Regression models in clinical studies: determining relationships between predictors and response. J. Natl. Cancer Inst. 80, 1198–1202, https://doi.org/10.1093/jnci/80.15.1198
41 Orsini, N., Bellocco, R. and Greenland, S. (2006) Generalized least squares for trend estimation of summarized dose-response data. Stat. J. 6, 40–57, https://doi.org/10.1177/153687290600600301
42 Jackson, D., White, I.R. and Thompson, S.G. (2010) Extending DerSimonian and Laird’s methodology to perform multivariate random effects meta-analyses. Stat. Med. 29, 1282–1297, https://doi.org/10.1002/sim.3602
43 Pike, M.C., Pearce, C.L., Peters, R., Cozen, W., Wan, P. and Wu, A.H. (2004) Hormonal factors and the risk of invasive ovarian cancer: a population-based case-control study. Fertil. Steril. 82, 186–195, https://doi.org/10.1016/j.fertnstert.2004.03.013
44 Koskela-Niska, V., Pukkala, E., Lyytinen, H., Ylikorkala, O. and Dyba, T. (2013) Effect of various forms of postmenopausal hormone therapy on the risk of ovarian cancer: a population-based case control study from Finland. Int. J. Cancer 133, 1680–1688, https://doi.org/10.1002/ijc.28167
45 Wu, A.H., Pearce, C.L., Lee, A.W., Tseng, C., Jotwani, A., Patel, P. et al. (2017) Timing of births and oral contraceptive use influences ovarian cancer risk. Int. J. Cancer 141, 2392–2399, https://doi.org/10.1002/ijc.30910
46 Bevier, M., Sundquist, J. and Hemminki, K. (2011) Does the time interval between first and last birth influence the risk of endometrial and ovarian cancer? Eur. J. Cancer 47, 586–591, https://doi.org/10.1016/j.ejca.2010.10.004
47 Yang, L., Klint, A., Lambe, M., Bellocco, R., Riman, T., Bergfeldt, K. et al. (2008) Predictors of ovarian cancer survival: a population-based prospective study in Sweden. Int. J. Cancer 123, 672–679, https://doi.org/10.1002/ijc.23429
48 Riman, T., Dickman, P.W., Nilsson, S., Correia, N., Nordlinder, H., Magnusson, C.M. et al. (2002) Risk factors for invasive epithelial ovarian cancer: results from a Swedish case-control study. Am. J. Epidemiol. 156, 363–373, https://doi.org/10.1093/aje/kwf048
49 Kvale, G., Heuch, I., Nilssen, S. and Beral, V. (1988) Reproductive factors and risk of ovarian cancer: a prospective study. Int. J. Cancer 42, 246–251, https://doi.org/10.1002/ijc.2910420217
50 Fleming, J.S., Beaugie, C.R., Hayv, I., Chenevix-Trench, G. and Tan, O.L. (2006) Incessant ovulation, inflammation and epithelial ovarian carcinogenesis: revisiting old hypotheses. Mol. Cell. Endocrinol. 247, 4–21, https://doi.org/10.1016/j.mce.2005.09.014
51 Fathalla, M.F. (1971) Incessant ovulation—a factor in ovarian neoplasia? Lancet2, 163, https://doi.org/10.1016/S0140-6736(71)92335-X
52 Cramer, D.W. and Welch, W.R. (1983) Determinants of ovarian cancer risk. II. Inferences regarding pathogenesis. J. Natl. Cancer Inst. 71, 717–721
53 O’Leary, P., Boyne, P., Flett, P., Belby, J. and James, I. (1991) Longitudinal assessment of changes in reproductive hormones during normal pregnancy. Clin. Chem. 37, 667–672
54 Bu, S.Z., Yin, D.L., Ren, X.H., Jiang, L.Z., Wu, Z.J., Gao, Q.R. et al. (1997) Progesterone induces apoptosis and up-regulation of p53 expression in human ovarian carcinoma cell lines. Cancer 79, 1944–1950, https://doi.org/10.1002/(SICI)1097-0142(19970515)79:10<1944::AID-CNCR15%3E3.0.CO;2-V
55 Adami, H.O., Hsieh, C.C., Lambe, M., Trichopoulos, D., Leon, D., Persson, I. et al. (1994) Parity, age at first childbirth, and risk of ovarian cancer. Lancet 344, 1250–1254, https://doi.org/10.1016/S0140-6736(84)90749-8
56 Pike, M.C. (1987) Age-related factors in cancers of the breast, ovary, and endometrium. J. Chronic Dis. 40, 59s–69s, https://doi.org/10.1016/S0021-9681(87)80009-7
57 Munafò, M.R. and Flint, J. (2004) Meta-analysis of genetic association studies. Trends Genet. 20, 439–444, https://doi.org/10.1016/j.tig.2004.06.014