Maternal age and risk of stillbirth: a systematic review

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Over the past several decades, economic, technologic and social changes in the developed world have significantly increased the number of women who delay childbirth to their late 30s and beyond. Between 1980 and 1993 in the European Union, the mean maternal age at first birth rose by 1.5 years, from 27.1 to 28.6 years.1 Between 1991 and 2001 in the United States, the percentage of first births for women 35–39 years of age increased by 36% and that for women 40–44 years of age increased by 70%.2,3 This huge demographic shift has become an important public health issue, since numerous studies have indicated that increased maternal age (35 years of age or older) is associated with an increased risk of maternal morbidity, obstetric interventions and adverse pregnancy outcomes.4–11

The intention to conceive and the timing of conception are complex issues influenced by many factors. Understanding the potential outcomes of pregnancy at an advanced maternal age may constitute one of these factors. Stillbirth, or late fetal death, is one of the adverse pregnancy outcomes of most concern, but studies on the relation between increased maternal age and stillbirth risk have led to inconsistent conclusions. We therefore conducted a systematic review of observational studies to explore the association between advanced maternal age and the risk of stillbirth.

Methods

Literature search

To identify relevant studies, we searched the medical literature for articles published up to Dec. 31, 2006, using MEDLINE (from 1966), EMBASE (from 1980) and the Cochrane Database of Systematic Reviews. The search terms included “maternal age,” “childbearing,” “fetal death,” “perinatal mortality” and terms identifying observational studies and systematic reviews (for search strategies, see online Appendix 1 available at www.cmaj.ca/cgi/content/full/178/2/165/DC2). We also hand-searched the bibliographies of retrieved articles to identify additional related studies. We did not attempt to identify articles from the grey literature (unpublished studies with limited distribution, e.g., conference proceedings, theses, website archives or newsletters). No language restriction was imposed.

Results:

We identified 913 unique citations, of which 31 retrospective cohort and 6 case–control studies met our inclusion criteria. In 24 (77%) of the 31 cohort studies and all 6 of the case–control studies, we found that greater maternal age was significantly associated with an increased risk of stillbirth; relative risks varied from 1.20 to 4.53 for older versus younger women. In the 14 studies that presented adjusted relative risk, we found no extensive change in the direction or magnitude of the relative risk after adjustment. We did not calculate a pooled relative risk because of the extreme methodologic heterogeneity among the individual studies.

Interpretation:

Women with advanced maternal age have an increased risk of stillbirth. However, the magnitude and mechanisms of the increased risk are not clear, and prospective studies are warranted.

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Study selection
We included all cohort studies (prospective or retrospective) and all case–control studies that examined the association between advanced maternal age and stillbirth risk. To be included in the systematic review, studies had to cite the odds ratio or relative risk of stillbirth by maternal age or had to provide sufficient data to allow us to construct contingency tables detailed enough to calculate the relative risk of stillbirth by maternal age.

Data abstraction and quality assessment
Two reviewers (L.H. and C.v.W.) independently abstracted data from all of the included studies. Abstracted data included study design; type of database used for analysis (population-based or hospital-based), if any; characteristics of the study subjects, including plurality (singletons, multiple births or both) and parity (nullipara, multipara or both); stillbirth definition in terms of gestational week and birth weight; types of stillbirths included; potential confounders or effect modifiers considered; numbers of live births and stillbirths by maternal age group; and risk ratio or odds ratio for stillbirth. We assessed the quality of each study using the Newcastle–Ottawa Scale for nonrandomized studies in meta-analyses. Differences in data abstraction were resolved by consensus after reference back to the original article. The interobserver agreement for data abstraction was measured using weighted kappa statistics.

Data analysis
We recorded both the crude and (if available) adjusted risk ratio or odds ratio and 95% confidence interval (CI) for stillbirth risk by maternal age for each individual study. Risk ratios and odds ratios greater than 1.0 indicate an increased risk of stillbirth among older women. We assessed the heterogeneity of the results of the individual studies using the Cochran Q-test for all studies, for population-based cohort studies, for hospital-based cohort studies and for case–control studies.

Results
Through the literature search, we initially identified 913 unique citations (Figure 1). After our detailed review, 37 studies (31 retrospective cohort studies and 6 case–control studies) met the inclusion criteria. Of the included studies, 16 were from Europe, 9 from North America, 5 from Asia, 4 from Latin America and 3 from Africa. Thirty-two of the studies were published in English, and the remaining 5 were published in French, Spanish, German or Italian. The literature search also identified 21 relevant systematic reviews. However, none of these reviews reported a systematic search for or analysis of individual studies focusing on the relation between advanced maternal age and stillbirth risk. Overall, the studies included in our review were of high quality: all studies scored 7 or higher on the Newcastle–Ottawa scale (maximum score 9). Of the 37 studies, 20 (54%) scored 9, 5 (14%) scored 8, and 12 (32%) scored 7 (Table 1). The kappa score was 0.92, which indicated very good agreement between the reviewers with regard to study quality score.

The characteristics of the individual studies are summarized in Appendix 2 (available online at www.cmaj.ca/cgi/content/full/178/2/165/DC2). Data collection extended from 1959 to as recently as 2003. Most of the studies were very large, with a median sample size of 78 453 births or women. In 17 of the studies, multiple births were included in the analysis. Studies varied with regard to criteria for defining stillbirth, and some had no specific gestational age or birth weight criteria for stillbirth. In 16 studies there was no adjustment for any other variables when the association of maternal age with stillbirth risk was measured. Studies varied extensively in their definitions of increased maternal age and in terms of the maternal age group used as the comparator. Fifteen studies excluded births among women younger than a particular age. Notably, in only 2 of the studies was maternal age entered into the regression model as a continuous variable, but these studies did not report a linearity check for the association between maternal age and stillbirth risk. It is important to take into account possible nonlinearity when a continuous exposure is included in conventional logistic regression models, since violation of the assumption of a linear logit relation between the continuous exposure and outcome can invalidate analyses.

We found statistical heterogeneity for all included studies, for the population-based cohort studies and for the hospital-based cohort studies ($p < 0.001$ for the 3 tests of heterogeneity). The results of the case–control studies were homogeneous ($p = 0.40$). Because of the statistical and important clinical and methodological heterogeneities, we did not perform meta-analytic combination of the results of these studies.
Among the 31 retrospective cohort studies, 24 (77%) had a statistically significant association between increased maternal age and stillbirth (Figure 2). In addition, for all 15 population-based studies, there was a significant association between increased maternal age and stillbirth risk. The reported crude relative risk of stillbirth varied from 1.26 to 4.53 for older versus younger women. The most commonly used definition of advanced maternal age in these studies was 35 years or more for older women, with age less than 35 years or 20–34 years used as the reference group (these 2 definitions were used in 10 of the 24 studies). The authors of these 10 studies reported a stillbirth risk among older women that was 1.26 to 1.92 times higher than the risk among women less than 35 years of age. In 7 of the 31 studies, there was no statistically significant relation between older maternal age and stillbirth risk. All of these 7 studies were based on hospital-level databases (Figure 2). In all 6 case–control studies, there was a significant relation between advanced maternal age and stillbirth risk; the crude odds ratios were similar to those reported for the cohort studies (Figure 2).

In 21 (57%) of the studies (15 retrospective cohort studies and all 6 case–control studies), the researchers controlled for potential confounders when measuring the association between maternal age and stillbirth risk. Both crude and adjusted risk ratios or odds ratios were reported for 14 of these 21 studies (Figure 3). Different confounders were adjusted for in each study, most frequently parity and smoking, followed by education, race, chronic medical problems, prenatal care and body mass index. After adjustment for confounders, the risk ratios or odds ratios did not change extensively in either magnitude or direction. The one exception was the study by Sheiner and colleagues, in which the crude relative risk of 1.48 (95% CI 0.74–2.96) among women more than 35 years of age increased to 2.10 (95% CI 1.10–3.80) after adjustment.

Interpretation

The number of women who are experiencing childbirth at an older age is increasing dramatically. We identified 37 studies that examined the association between increased maternal age and stillbirth, more than 80% of which demonstrated a statistically significant increase in the risk

Table 1: Quality assessment* of 37 individual studies of the association between advanced maternal age and stillbirth risk that were included in the systematic review

| Study                          | Selection | Comparability | Outcome or exposure | Total score |
|-------------------------------|-----------|---------------|---------------------|-------------|
| **Retrospective cohort studies (n = 31)** |           |               |                     |             |
| Astolfi et al14               | ★★★★★    | ★★            | ★★★★               | 9           |
| Astolfi et al15               | ★★★★★    | ★★            | ★★★★               | 9           |
| Bianco et al16                | ★★★★★    | ★★            | ★★★★               | 9           |
| Canterino et al17             | ★★★★★    | ★★            | ★★★★               | 9           |
| Cnattingius et al11           | ★★★★★    | ★★            | ★★★★               | 9           |
| Conde-Agudelo et al18         | ★★★★★    | ★★            | ★★★★               | 9           |
| Donoso et al19                | ★★★★★    | ★★            | ★★★★               | 7           |
| Feldman20                     | ★★★★★    | ★★            | ★★★★               | 8           |
| Feresu et al21                | ★★★★★    | ★★            | ★★★★               | 7           |
| Fretts et al10                | ★★★★★    | ★★            | ★★★★               | 9           |
| Fretts et al22                | ★★★★★    | ★★            | ★★★★               | 7           |
| Gadow et al21                 | ★★★★★    | ★★            | ★★★★               | 7           |
| Glinianaia et al24            | ★★★★★    | ★★            | ★★★★               | 8           |
| Haglund et al25               | ★★★★★    | ★★            | ★★★★               | 9           |
| Heimann et al26               | ★★★★★    | ★★            | ★★★★               | 7           |
| Jacobsson et al27             | ★★★★★    | ★★            | ★★★★               | 9           |
| Jolly et al7                  | ★★★★★    | ★★            | ★★★★               | 9           |
| Khandait et al28              | ★★★★★    | ★★            | ★★★★               | 7           |
| Kristensen et al29            | ★★★★★    | ★★            | ★★★★               | 9           |
| Lammer et al10                | ★★★★★    | ★★            | ★★★★               | 7           |
| Naeye31                       | ★★★★★    | ★★            | ★★★★               | 7           |
| Nybo Anderson et al32         | ★★★★★    | ★★            | ★★★★               | 7           |
| Pugliese et al33              | ★★★★★    | ★★            | ★★★★               | 7           |
| Rasmussen et al34             | ★★★★★    | ★★            | ★★★★               | 8           |
| Raymond et al35               | ★★★★★    | ★★            | ★★★★               | 9           |
| Roman et al36                 | ★★★★★    | ★★            | ★★★★               | 8           |
| Seoud et al37                 | ★★★★★    | ★★            | ★★★★               | 8           |
| Sheiner et al38               | ★★★★★    | ★★            | ★★★★               | 9           |
| Tough et al39                 | ★★★★★    | ★★            | ★★★★               | 7           |
| Viegas et al40                | ★★★★★    | ★★            | ★★★★               | 9           |
| Ziadeh41                      | ★★★★★    | ★★            | ★★★★               | 7           |
| **Case-control studies (n = 6)** |           |               |                     |             |
| Ferraz et al42                | ★★★★★    | ★★            | ★★★★               | 9           |
| Little et al43                | ★★★★★    | ★★            | ★★★★               | 9           |
| Meda et al44                  | ★★★★★    | ★★            | ★★★★               | 9           |
| Petridou et al45              | ★★★★★    | ★★            | ★★★★               | 9           |
| Smeeton et al46               | ★★★★★    | ★★            | ★★★★               | 9           |
| Stephansson et al47           | ★★★★★    | ★★            | ★★★★               | 9           |

*According to the Newcastle-Ottawa Scale for nonrandomized studies in meta-analyses. Maximum score for selection = 4 stars, maximum score for comparability = 2 stars, maximum score for outcome (for retrospective cohort studies) or exposure (for case-control studies) = 3 stars, maximum total score = 9.
Table 1: Crude risk of stillbirth among older and younger women in 31 retrospective cohort studies and 6 case–control studies. The cutoff for the 2 age groups (younger v. older) was variable. OR = odds ratio, RR = risk ratio, CI = confidence interval.

| Study                  | Older women (n/N)* | Younger women (n/N)* | OR or RR (95% CI)       |
|------------------------|--------------------|----------------------|------------------------|
| **Population-based**   |                    |                      |                        |
| Astolfi et al46        | 1 174/202 054      | 8 455/2 089 070      | 1.44 (1.35-1.53)       |
| Astolfi et al46        | 3 228/512 381      | 7 932/1 983 000      | 1.57 (1.51-1.64)       |
| Cantrero et al47       | 8 177/2 470 005    | 15 061/5 440 685     | 1.20 (1.16-1.23)       |
| Cnattingus et al40     | 41/7 880           | 251/70 397           | 1.46 (1.05-2.03)       |
| Conde-Aguado et al46   | 2 675/87 063       | 9 400/579 389        | 1.89 (1.82-1.98)       |
| Donoso et al47         | 6 657/1 066        | 742/175 256          | 2.18 (1.69-2.82)       |
| Feldman et al32        | 441/46 596         | 2 013/322 455        | 1.53 (1.38-1.69)       |
| Gliniana et al44       | 365/53 448         | 973/228 694          | 1.61 (1.42-1.81)       |
| Haglund et al40        | 95/21 467          | 558/163 689          | 1.30 (1.04-1.61)       |
| Jacobson et al47       | 217/32 867         | 2 785/876 361        | 2.08 (1.81-2.38)       |
| Lammer et al47         | 59/5 086           | 435/63 044           | 1.69 (1.29-2.22)       |
| Nybo Anderson et al48  | 261/48 950         | 3 238/763 411        | 1.26 (1.11-1.43)       |
| Rasmussen et al47      | 457/163 782        | 130/83 820           | 1.80 (1.46-2.19)       |
| Raymond et al47        | 349/75 062         | 1 811/563 180        | 1.45 (1.29-1.62)       |
| Tough et al47          | 270/29 233         | 1 643/254 723        | 1.44 (1.26-1.63)       |
| **Hospital-based**     |                    |                      |                        |
| Bianco et al46         | 2/1 404            | 29/6 978             | 0.34 (0.08-1.43)       |
| Fersu et al48          | 102/1 177          | 700/12 333           | 1.53 (1.25-1.86)       |
| Frett et al49          | 127/11 942         | 561/82 376           | 1.57 (1.29-1.90)       |
| Frett et al49          | 133/13 222         | 582/88 390           | 1.53 (1.27-1.85)       |
| Gadow et al45          | 3 000/91 416       | 12 937/756 803       | 1.92 (1.85-2.00)       |
| Heimann et al46        | 7/560              | 11/615               | 0.70 (0.27-1.81)       |
| Jolly et al48          | 312/48 658         | 1 581/336 462        | 1.37 (1.21-1.54)       |
| Khandait et al49       | 160/2 700          | 920/41 913           | 2.81 (2.36-3.34)       |
| Kristensen et al48     | 20/3 163           | 92/21 342            | 1.47 (0.91-2.38)       |
| Naeye et al49          | 110/3 346          | 605/41 040           | 2.23 (1.83-2.72)       |
| Puglise et al49        | 1/314              | 60/6 683             | 0.35 (0.05-2.55)       |
| Romano et al49         | 10/782             | 56/7 786             | 3.64 (1.87-7.08)       |
| Seoud et al49          | 17/319             | 4/326                | 4.53 (4.51-18.68)      |
| Sheiner et al49        | 10/10 114          | 41/61 463            | 1.48 (0.74-2.96)       |
| Viegas et al49         | 10/2 410           | 101/19 032           | 0.78 (0.41-1.50)       |
| Ziadeh et al49         | 3/143              | 1/148                | 3.15 (0.25-166.47)     |

*For each group, n = number of women experiencing stillbirth (number of cases in case–control studies), and N = total number of women in the sample (number of cases plus controls in case–control studies).
| Retrospective cohort studies | Variables adjusted for                                                                 | Age stratification, yr | OR or RR (95% CI)               |
|-----------------------------|----------------------------------------------------------------------------------------|------------------------|---------------------------------|
| Astone et al15               | Paternal age, parity and family education                                               | ≥35 v. 20-29           | Unadjusted 1.57 (1.51–1.64)     |
|                             |                                                                                       |                        | Adjusted 1.59 (1.51–1.68)       |
| Bianco et al16              | Chronic medical illness, race and tobacco use                                           | ≥40 v. 20-29           | Unadjusted 0.34 (0.08–1.43)     |
|                             |                                                                                       |                        | Adjusted 0.50 (0.10–2.00)       |
| Canterino et al17           | Gravidaity, race, marital status, prenatal care, smoking, education, abruptio placentae and birth cohort | 35-39 v. 20-24         | Unadjusted 1.11 (1.07–1.14)     |
|                             |                                                                                       |                        | Adjusted 1.23 (1.17–1.30)       |
|                             |                                                                                       | 40-44 v. 20-24         | Unadjusted 1.57 (1.50–1.65)     |
|                             |                                                                                       |                        | Adjusted 1.62 (1.48–1.76)       |
|                             |                                                                                       | 45-49 v. 20-24         | Unadjusted 2.59 (2.17–3.10)     |
|                             |                                                                                       |                        | Adjusted 2.40 (1.77–3.27)       |
| Cnattingius et al11         | Hypertension, diabetes, antepartum hemorrhage, infertility, education, cohabitation with infant’s father and smoking | ≥35 v. 20-24           | Unadjusted 1.46 (1.05–2.03)     |
|                             |                                                                                       |                        | Adjusted 1.40 (1.00–2.00)       |
| Conde-Agudelo et al18       | Education, antenatal care, body mass index, parity, small for gestational age, previous stillbirth and medical conditions | ≥35 v. 20-34           | Unadjusted 1.89 (1.82–1.98)     |
|                             |                                                                                       |                        | Adjusted 1.53 (1.42–1.55)       |
| Feresu et al21              | Prenatal care, parity, fetal sex, type of delivery and residence                       | > 35 v. 20-35          | Unadjusted 1.53 (1.25–1.86)     |
|                             |                                                                                       |                        | Adjusted 1.46 (1.12–1.89)       |
| Jacobsson et al37           | Nulliparity, marital status, maternal diseases, smoking, plurality and malformations   | 40-44 v. 20-34         | Unadjusted 2.02 (1.75–2.33)     |
|                             |                                                                                       |                        | Adjusted 2.07 (1.79–2.40)       |
|                             |                                                                                       | ≥45 v. 20-34           | Unadjusted 3.66 (2.17–6.16)     |
|                             |                                                                                       |                        | Adjusted 3.76 (2.22–6.40)       |
| Jolly et al7                | Parity, race, body mass index, hypertension, pre-eclampsia, diabetes and smoking       | 35-40 v. 18-34         | Unadjusted 1.37 (1.21–1.54)     |
|                             |                                                                                       |                        | Adjusted 1.41 (1.17–1.70)       |
|                             |                                                                                       | > 40 v. 18-34          | Unadjusted 1.72 (1.32–2.23)     |
|                             |                                                                                       |                        | Adjusted 1.83 (1.29–2.61)       |
| Rasmussen et al44           | Birth order                                                                            | 35-39 v. < 20          | Unadjusted 1.69 (1.38–2.07)     |
|                             |                                                                                       |                        | Adjusted 1.70 (1.40–2.10)       |
|                             |                                                                                       | 40-44 v. < 20          | Unadjusted 2.13 (1.63–2.78)     |
|                             |                                                                                       |                        | Adjusted 2.00 (1.50–2.60)       |
|                             |                                                                                       | ≥45 v. < 20            | Unadjusted 4.81 (2.72–8.53)     |
|                             |                                                                                       |                        | Adjusted 3.90 (2.20–7.00)       |
| Raymond et al45             | Parity and smoking                                                                     | ≥35 v. 20-34           | Unadjusted 1.45 (1.29–1.62)     |
|                             |                                                                                       |                        | Adjusted 1.50 (1.40–1.70)       |
| Sheiner et al48             | Birth weight, race, abruptio placentae, congenital anomalies, cord problem, polyhydramnos and malpresentation | > 35 v. 20-35         | Unadjusted 1.48 (0.74–2.96)     |
|                             |                                                                                       |                        | Adjusted 2.10 (1.10–3.80)       |
| Case-control studies        |                                                                                       |                        |                                    |
| Little et al44              | Education, race, body mass index, drinking, smoking, maternal height and weight, region of birth, fetal sex and nulliparity | ≥35 v. < 35           | Unadjusted 1.61 (1.19–2.18)     |
|                             |                                                                                       |                        | Adjusted 1.67 (1.27–2.20)       |
| Petridou et al46            | Fetal sex, plurality, gestational age, birth weight, education parity, religion and maternal residence | 35-39 v. 20-24         | Unadjusted 1.79 (1.50–2.14)     |
|                             |                                                                                       |                        | Adjusted 1.32 (1.05–1.65)       |
|                             |                                                                                       | ≥40 v. 20-24           | Unadjusted 2.30 (1.72–3.08)     |
|                             |                                                                                       |                        | Adjusted 2.15 (1.51–3.06)       |
| Smeeton et al46             | Gestational age, birth weight and emergency cesarean section                            | ≥30 v. < 30            | Unadjusted 1.55 (1.20–1.99)     |
|                             |                                                                                       |                        | Adjusted 1.07 (1.02–1.12)*      |

*Per year increase.

Figure 3: Crude and adjusted risk of stillbirth in the 11 retrospective cohort studies and 3 case–control studies that reported an adjusted odds ratio (OR) or risk ratio (RR). For the last case–control study (Smeeton et al46), the adjusted OR refers to the per-year increase in risk of stillbirth. Plurality refers to singleton v. multiple birth. CI = confidence interval.
of stillbirth among women of advanced maternal age. However, the absolute increase in risk was relatively small in studies from developed countries, with crude odds ratios varying from 1.20 to 2.23 on top of baseline stillbirth rates varying from 1.55 to 17.89 per 1000 total births. Fourteen studies adjusted this association for potential confounders, and 8 of them found less than a 10% difference between crude and adjusted relative risks or odds ratios. This result suggests that the association between increased maternal age and stillbirth risk is probably independent of many other known factors that could increase with age.

The biological mechanism of the increase in stillbirth risk with advanced maternal age is uncertain. A direct effect of maternal aging may exist. This would probably be related to low uteroplacental perfusion caused by poor uterine vasculature in older women. The increased risk could also be attributed to the association between older age and certain risk factors for stillbirths, such as chronic diseases and medical or obstetric complications. Older women have a higher risk of experiencing pregnancy-induced hypertension or gestational diabetes. Between 50% and 70% of mothers of stillborn infants had medical or pregnancy complications during their pregnancies. The relative contributions of advanced maternal age and chronic conditions that may influence pregnancy outcomes remain to be fully elucidated. Additional studies are required to determine the mechanism by which advanced maternal age increases stillbirth risk.

Maternal age and parity are 2 closely related demographic factors. They are commonly included in obstetric care records because of their influences on pregnancy complications and outcomes. In clinical practice, both older nulliparous and younger multiparous women are considered to be at increased risk of adverse pregnancy outcomes. Studies have demonstrated that the association of maternal age with outcomes, including placenta previa and abruptio placentae and neonatal mortality, varies significantly by parity. However, few researchers have used multivariate regression analysis to address the interaction between parity and maternal age in terms of its effect on the risk of stillbirths. For example, only 3 (8%) of the 37 individual studies in our review included a test for interaction between parity and maternal age. Further study is required to explore this interaction in more detail.

We did not perform a meta-analysis to determine an overall effect measure because of extensive heterogeneity in study design, measurement and even definition of stillbirth across the studies. There were several other notable variations among the studies. First, the definition of “increased maternal age” was variable (see Appendix 2, available online at www.cmaj.ca/cgi/content/full/178/2/165/DC2). This variation probably reflects the clinical uncertainty about the maternal age at which stillbirth risk increases. In 1958 the International Federation of Gynecology and Obstetrics defined pregnant women 35 years or older as “elderly primigravidae.” In obstetric practice, pregnant women older than 35 years have been regarded as being at higher risk and are considered for routine genetic screening. However, the 35-year threshold is not applicable for healthy women who do not engage in risky behaviours such as smoking, because neonatal outcomes in such women do not change up to the age of 40 years. Other researchers have even suggested that, in the absence of pre-existing medical disorders, pregnancy at age 45 to 50 years of age is safe. Because it is difficult to assign a risk cut point to a continuous variable such as age, we believe that future studies should model age as a continuous variable when its association with stillbirth risk is being considered.

The inclusion or exclusion of multiple births was a second major source of heterogeneity among the studies, with some researchers including all deliveries and others excluding women with multiple births. Complications such as low birth weight, preterm labour and stillbirth are more likely with multiple births. Older women are 1.5 to 2 times more likely than younger women to have a multiple birth, at least in part because of their greater use of assisted reproductive technology. Therefore, we recommend that future studies examining the association between increased maternal age and stillbirth risk exclude multiple births or carefully consider their effect.

Study setting was a third major source of heterogeneity. Of the 16 hospital-based cohort studies, 7 had no statistically significant increase in the risk of stillbirth with older maternal age. Most of these 7 studies were performed in academic hospitals or medical centres. Older women with high socioeconomic status may seek obstetric care in academic hospitals, especially when delivering their first baby. At the same time, many of these hospitals are located in impoverished neighbourhoods, where young women with low socioeconomic status typically receive their routine obstetric care. An increased risk of adverse obstetric outcomes for these women may obscure the association between advanced maternal age and stillbirth risk.

The association between increased maternal age and stillbirth risk was essentially similar across the studies, despite the extensive methodologic heterogeneity. In particular, population-based studies with large sample sizes, which accounted for almost half of the studies included in our review, had consistent results regarding this association. Also, adjustment for a wide variety of potential confounders had little effect on the association between maternal age and stillbirth (Figure 3). We believe that these observations suggest a stable and valid association between increased maternal age and stillbirth risk.

Women should understand that the risks associated with pregnancy increase as they get older. In this systematic review, we have demonstrated that risk of stillbirth is one of the risks that increases with advanced maternal age. However, since the absolute stillbirth rate among older women is less than 10 per 1000 births in most industrialized countries, a live birth can be expected in most cases if appropriate medical care is provided. Thus, increased maternal age should not be considered an absolute barrier to the decision to have a child, at least in terms of the risk of stillbirth.

Limitations

The limitations of this systematic review stem primarily from the studies themselves. First, we did not perform a meta-analysis or meta-regression because of the extreme method-
oligoc heterogeneity across the studies. Second, few of the included studies reported a separate analysis for antepartum stillbirths (which occur before labour) and intrapartum stillbirths (which occur during labour). Given that these 2 types of stillbirth have widely different etiologic determinants, merging them might cloud studies of risk factors. Third, because information on important factors such as medical conditions, pregnancy complications, lifestyle factors and socioeconomic status are not commonly collected in administrative databases, many studies could not adjust for these factors when associating maternal age and stillbirth risk, which might weaken the internal validity of the included studies. However, investigators should be cautious when controlling for factors such as medical conditions or low birth weight in examining the relation between prenatal exposure and perinatal outcome, since these factors might be part of a causal pathway between the exposure and the outcome. In a recent study, Hernandez-Diaz and associates concluded that adjusting for birth weight is generally inappropriate when measuring the overall effect of maternal smoking on infant mortality; this conclusion arose through a paradoxical finding of the birth-weight effect, whereby small babies in high-risk populations usually have lower risk than small babies in low-risk populations.

Implications

In this systematic review, we have demonstrated that women with advanced maternal age have a significantly increased risk of stillbirth. We could not determine a pooled estimate of the magnitude of this risk because of heterogeneity in terms of the subjects included in the individual studies, the definitions of stillbirth, the cutoff point for older maternal age and the confounders that were adjusted for across the included studies. More studies with homogeneous design are needed to further evaluate the strength of the association between advanced maternal age and increased risk of stillbirth. Future studies should adopt the definition of stillbirth recommended by the World Health Organization to make comparisons between studies both feasible and accurate. Since the pathways through which older age influences the risk of stillbirth are still unclear, future studies should examine the mechanisms by which advanced maternal age leads to increased risk of stillbirth.

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