Heterogeneity of antidiabetic treatment effect on the risk of major adverse cardiovascular events in type 2 diabetes: a systematic review and meta-analysis

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Abstract

Background We explored whether clinically relevant baseline characteristics of patients with type 2 diabetes can modify the effect of glucagon-like peptide-1 receptor agonists (GLP-1 RA) or sodium-glucose cotransporter-2 inhibitors (SGLT-2i) on the risk of major adverse cardiovascular events (MACE).

Methods We investigated Medline and EMBASE through June 2019. We included randomized clinical trials reporting the effect of GLP-1 RA or SGLT-2i on MACE in subgroups of patients with type 2 diabetes, identified through key baseline factors: established cardiovascular disease; heart failure; chronic kidney disease; uncontrolled diabetes; duration of diabetes; hypertension; obesity; age; gender and race. Hazard ratios (HRs) and 95% confidence intervals (CIs) from trials were meta-analyzed using random-effects models.

Results Nine trials enrolling 87,143 patients were included in the analyses. Subgroup meta-analyses showed a 14% risk reduction of MACE in patients with established cardiovascular disease [GLP1-RA: HR, 0.86 (95% CI, 0.79-0.94); SGLT-2i: 0.86 (0.80-0.93)], and no effect in at-risk patients without history of cardiovascular events [GLP1-RA: 0.95 (0.83-1.08); SGLT-2i: 1.00 (0.87-1.16)]. We observed a trend toward larger treatment benefits with SGLT-2i among patients with chronic kidney disease [0.82 (0.69-0.97)], and patients with uncontrolled diabetes for both GLP1-RA or SGLT-2i [GLP1-RA: 0.82 (0.71-0.95); SGLT-2i: 0.84 (0.75-0.95)]. Uncontrolled hypertension, obesity, gender, age and race did not appear to modify the effect of these drugs.

Conclusions In this exploratory analysis, history of cardiovascular disease appeared to modify the treatment effect of SGLT2i or GLP1-RA on MACE. Chronic kidney disease and uncontrolled diabetes should be further investigated as potential effect modifiers.

Introduction

Type 2 diabetes mellitus (T2DM) is the seventh-leading cause of death in the US and a major contributor to cardiovascular disease [1]. Its increasing prevalence has translated to a commensurate growth of T2DM-related mortality and complications in recent years [1], which increases the urgency of implementing favorable findings from trials on anti-diabetic treatments into clinical practice.

A major barrier to effective implementation of such findings is the scarcity of evidence describing the extent to which results from trials may be generalizable to all patients with T2DM [2] or whether the results may vary across subgroups of the population [3, 4]. Exploring treatment effect heterogeneity in a trial population may provide useful information that could guide clinical decision making on which groups of patients may optimally benefit from a specific therapeutic strategy [5].

The identification of subgroups of interest in clinical trials usually emerges from baseline characteristics of the population, which have been deemed a priori as potential treatment effect modifiers [5, 6]. However, subgroup analyses within a trial are generally under-powered to detect treatment effect heterogeneity [5, 6].

In cardiovascular outcome trials mandated by FDA [7], glucagon-like peptide-1 receptor agonists (GLP-1 RA) and sodium-glucose cotransporter-2 inhibitors (SGLT-2i) have shown a reduction in the risk of major adverse cardiovascular events compared to placebo [8, 9]. However, less is known about potential effect modifiers of such treatment effects, and consequently, uncertainty remains on the patient populations who might gain higher benefits from these therapies in practice. Thus, we conducted an exploratory meta-analysis of placebo-controlled randomized controlled trials to examine whether baseline characteristics—identified as potential effect modifiers in the pre-defined statistical analysis of these trials—appeared to modify the effect of SGLT-2i and GLP-1 RA drugs on the outcome of major cardiovascular events.

Materials And Methods

Data Sources and Searches

Findings are reported following the Preferred Reporting Items for Systematic Review and Meta-analyses guidelines [10]. Two authors (E.D. and E.J.) investigated Medline and EMBASE from inception to August 2020 using search terms developed to cover
relevant drug classes and agents, type 2 diabetes, and cardiovascular outcomes (see eTable 1 in the Supplement). Reference lists of original articles, systematic reviews and meta-analyses were also screened.

**Study Selection**

Studies were required to contain the following inclusion criteria to be eligible for the meta-analyses: (1) report on double-blinded randomized controlled trials; (2) include no active comparisons (control arm should have received placebo or no active treatment); (3) report major adverse cardiovascular events (MACE) as primary outcome; (4) enroll participants with type 2 diabetes; (5) follow-up patients for longer than 6 months; (6) describe phase 3 trial dosage; (7) investigate products within the SGLT-2i and GLP-1 RA drug classes. We excluded trials if they did not test the effect of the intervention on MACE in predefined subgroups reported in the pre-specified statistical plan and if their results were not available in peer-reviewed articles.

The potential modifiers included in the current study were baseline factors of the trial populations measured at randomization, identified *a priori* based on their clinical relevance in a type 2 diabetes population. We classified these factors in four groups: cardiovascular factors (established atherosclerotic cardiovascular disease and heart failure), renal function (estimated glomerular filtration rate [eGFR], as indicator of chronic kidney disease [CKD]), cardiometabolic factors (HbA1c, duration of diabetes, BMI and systolic [SBP] and diastolic blood pressure [DBP]) and demographic factors (age, gender, race). The definitions and categorizations of the modifier across the trials are detailed in eTable 2 in the Supplement. All potential effect modifiers considered were part of the pre-specified subgroup analyses of the included trials, and findings were provided in the published articles.

**Data Extraction and Quality Assessment**

The following information was then independently extracted from each trial by two investigators (E.D. and E.J.): authors, year of publication, experimental and comparison drug, trial duration, length of follow-up, number of centers and countries, number of patients per trial arm, age, diabetes duration, gender prevalence, history of cardiovascular disease, body mass index (BMI), glycated hemoglobin (HbA1c) values, and hazard ratios (HRs) and 95% confidence intervals (CIs) for the treatment effect on major cardiovascular events in the overall population and in the subgroups evaluated by the trials. If trials failed to report exact HRs and 95% CIs for the subgroup analyses in the text, E.J. and E.D. extracted those data from graphs using different software (Get Data® and WebPlotDigitizer®) and comparing the obtained estimates. Specifically, we extracted from graphs values for the subgroups of the ELIXA trial and for some subgroups of the DECLARE trials (i.e., HbA1c, duration of diabetes, BMI, hypertension, gender, race, age). Discrepancies were resolved by consensus. The results on methodological quality are presented in the Supplement as risk of bias table [12].

**Data Synthesis and Analysis**

First, we conducted a random-effects meta-analysis [12] of the overall efficacy of SGLT-2i and GLP-1 RA drugs in reducing MACE outcomes in the individual trials. Second, we performed random-effects meta-analyses [13] of the efficacy of these drugs on MACE outcomes within the pre-specified subgroups reported in the included trials and listed above.

Subgroup meta-analyses were conducted among subgroups of individuals with the following specific characteristics at randomization: established atherosclerotic cardiovascular disease versus cardiovascular disease without events at randomization; history of heart failure or congestive heart failure versus not; presence of chronic kidney disease (defined as eGFR levels less than 60 mL/min/1.73m²) versus not; uncontrolled diabetes, defined as HbA1c higher than 8%, versus better controlled diabetes, defined as HbA1c lower than 8% (8% was the most common HbA1c target reported in the pre-specified subgroups of the trials to compare controlled vs uncontrolled diabetes, we did not include trials that reported a threshold less than 8% and equal or higher than 8.5%); duration of diabetes (longer versus shorter than 10 years); hypertension (SBP ≥ 130 mmHg and/or DBP ≥ 80 mmHg) versus not; obesity (BMI ≥ 30.0) versus not; age equal to 65 years and younger versus older than 65 years; gender (male vs female) and race (white vs black vs Asian).

When a pre-specified subgroup analysis in a trial was conducted across more than two levels of the modifier, we used fixed-effect model [14] to pool the estimates across subgroups and create the comparisons listed above (for example if separate results were...
reported for the subgroup “age between 65 and 75 years” and “age greater than 75 years,” these were pooled to produce a subgroup of “age greater than 65” and compare it with the subgroup “age lower than or equal to 65 years”.

In a few cases the pre-specified subgroups from the trials did not meet exactly our subgroup definitions (i.e., the LEADER trial used 60 years old as a cut-off for age; REWIND trial defined obesity using a BMI equal to 32 kg/m2). We included these subgroups in the main meta-analyses (i.e., for the LEADER trial, patients between 61 and 65 years old were included in the subgroup of patients older than 65; for the REWIND trial, patients with a BMI equal to 31 and 32 kg/m2 were assigned to the subgroup of patients with BMI ≤ 30 kg/m2), under the hypothesis that these discrepancies would not affect the summarized results. We then conducted sensitivity analyses excluding these trials to test the robustness of our main summarized results.

We performed a random-effects meta-regression, using the restricted maximum likelihood estimator with Knapp-Hartung modification, to assess the differences in the treatment effect by drug class [15]. When there was not significant between-drug class difference of the treatment effect, we combined the two drug classes and performed a test to evaluate differences between subgroups, using the Cochran’s Q test [16]. Results are presented as HRs with 95% CIs. Between-study heterogeneity was assessed using the I² statistic [17]. However, when data are limited, I² and 95% CI are typically large, and magnitude of the statistical heterogeneity (conventionally described as low for I² values between 25%–50%, moderate for 50%–75%, and high for ≥75% [17]) should be interpreted with caution [18]. Funnel plot and Egger’s tests were conducted to estimate potential selection biases [19], and results are reported in the Supplement.

All analyses conducted to evaluate potential treatment heterogeneity across subgroups were exploratory. Therefore, p values were not adjusted according to the number of comparisons [20] and were regarded as significant when lower than 0.05. STATA version 15.0 was used for all calculations (College Station, Texas, Stata Corporation, 2017).

Results

The literature search identified 5,809 articles (eFigure 1 in the Supplement), of which 10 trials met the inclusion criteria [11, 21-29]. 89,790 patients were enrolled, including 34,322 for SGLT-2is and 55,438 for GLP-1 RAs. All studies presented an overall low risk of bias and there was no of publication bias (eTable 3-4 and eFigure 2 in the Supplement). Baseline information is shown in Table 1. Mean age across the entire sample was 63.5 years (range: 60.3-66.2), and mean BMI was 31.9 (range: 30.2-32.8), duration of diabetes ranged from 9.3 to 14.9 years. The median (interquartile range) duration of follow-up was 2.9 (2.2 - 3.7) years. The studies enrolled mostly men (range: 54-71%). Three studies exclusively enrolled participants with established cardiovascular disease [21, 25, 26]. The comparator was placebo in all studies. In nine trials, the primary outcome was a composite of cardiovascular death, nonfatal myocardial infarction, and nonfatal stroke, i.e., three-point MACE (3P-MACE) [11, 22-29], while one trial added hospitalization for unstable angina to these three endpoints (4P-MACE) [21]. All drugs tested in the trials have been approved by the Food and Drug Administration (FDA) as a therapeutic option for the treatment of type 2 diabetes.

Efficacy of GLP1-RA and SGLT-2i on MACE

Overall, SGLT-2i and GLP1-RA showed 11% [HR 0.89, (95% CI 0.83-0.96)] and 12% [HR 0.88, (95% CI 0.82-0.94)] reduction of major cardiovascular events, respectively (Figure 1).

Seven trials reported a subgroup analysis by history of cardiovascular disease [11, 22-24, 27-29], while another three included only patients who experienced at least one cardiovascular event [21, 25, 26]. In 63,105 patients with established cardiovascular disease, GLP1-RA and SGLT-2i drugs showed a 14% reduction of MACE [GLP1-RA: 0.86 (0.80-0.93); SGLT-2i: 0.86 (0.80-0.93)]. By contrast, in 26,665 patients at high risk of cardiovascular disease, but without history of cardiovascular events, GLP1-RA and SGLT-2i seemed to have minimal or no effect on MACE [GLP1-RA: 0.94 (0.82-1.07); SGLT-2i: 1.00 (0.87-1.16)] (difference in effect between patients with vs. without a history of cardiovascular disease: p = 0.049, I² = 74%) (Figure 2).

Seven trials reported results stratified by prior heart failure [20-24, 26, 27]. SGLT-2i and GLP1-RA drugs showed a risk reduction in MACE of 10% and 14% in the subgroups of patients without prior heart failure [HR 0.90 (0.83-0.98), n patients = 3,185, and HR 0.86...
(0.78-0.96), n patients = 7,497], and a 9-10% risk reduction among patients with prior heart failure [0.91 (0.73-1.14), n patients = 24,117, and 0.90 (0.79-1.02), n patients = 35,372], respectively (difference in effect between patients with vs. without heart failure: \( p = 0.652 \)) (eFigure 3 in the Supplement).

A subgroup analysis by eGFR levels was reported in nine trials [21-29]. The number of patients with impaired renal function was about one sixth of the number of patients with regular or mild impaired renal function in the SGLT-2i trials (n patients = 29,196 vs 5,123) and approximately one third in the GLP1-RA trials (n patients = 35,251 vs 10,773). Compared to placebo, SGLT2i drugs showed a trend towards larger reduction in the risk of MACE among patients with CKD than among patients without CKD [0.82 (0.69-0.97) vs 0.91 (0.83-1.00)], (\( p = 0.307 \)). GLP-1 RA drugs appeared to have similar reductions in the risk of MACE, independently of history of CKD [patients with CKD: 0.88 (0.75-1.03) vs patients without CKD: 0.85 (0.75-0.97)] (Figure 3).

All trials reported findings stratified by HbA1c level. Five trials selected 8% as the threshold to identify patient subgroups [22, 24, 25, 27, 28]. In patients with uncontrolled diabetes (HbA1c>8%), SGLT2i and GLP1-RA drugs reduced the risk of MACE by 16% [0.84 (0.75-0.95)] and 18% [0.82 (0.71-0.95)], respectively; while in patients with a better diabetes control (HbA1c≤8%), the risk reduction was 8-9% [GLP1-RA: 0.91 (0.82-1.00); SGLT-2i: 0.92 (0.79-1.07)] (Figure 4, \( p = 0.152 \)). Duration of diabetes did not appear to modify the effect of GLP1-RA drugs on MACE across subgroups [duration < 10 years: 0.85 (0.72-1.01); duration \( \geq 10 \) years: 0.88 (0.82-0.95)]; SGLT-2i drugs showed a 14% reduction in the risk of MACE in patients with a history of diabetes longer than ten years and had null effect on those with a shorter one [duration <10 years: 0.86 (0.79-0.93) ; duration \( \geq 10 \) years: 1.03 (0.94-1.13)] (difference between patients with diabetes < 10 years vs. duration \( \geq 10 \) years: \( p = 0.472 \)) (eFigure 4 in the Supplement). The effect of GLP1-RA and SGLT-2i drugs on MACE appeared to be similar in groups of patients with or without obesity (difference between patients with vs. without obesity: \( p = 0.789 \)) (eFigure 5 in the Supplement). Sensitivity analyses excluding REWIND trial from the GLP1 RA trials yielded similar effect estimates and confidence intervals compared to the main analyses stratified by BMI (eFigure 10 in the Supplement).

Similarly, SGLT-2i drugs showed equal risk reductions in subgroups of patients with and without hypertensionfor the three trials that included the subgroup analyses for uncontrolled hypertension [26-28] (eFigure 6 in the Supplement, \( p = 0.924 \)).

All included trials explored the effect of gender, race, and age on MACE [11,21-29]. The effect of GLP1-RA and SGLT-2i drugs on MACE was similar across all these factors (e.g., difference between male vs female patients, \( p = 0.375 \)), though we observed a trend towards larger reductions in the risk of MACE for GLP1-RA among Asian patients [0.72 (0.58-0.90)], compared with white and black individuals [0.89 (0.80-0.99) and 1.02 (0.54-1.62)] (overall difference between white vs. Asian patients: \( p = 0.085 \)), and for SGLT-2i among individuals aged 65 years or older [0.81 (0.69-0.96)], compared to those younger [0.95 (0.85-1.07)] (difference between patients younger vs. older than 65 years old: \( p = 0.585 \)) (eFigure 7-9 in the Supplement). Sensitivity analyses excluding the LEADER trial from the GLP1 RA trials yielded almost identical effect estimates and confidence intervals of the main subgroup meta-analyses stratified by age (eFigure 11).

**Discussion**

This exploratory meta-analysis suggests that established cardiovascular disease might be a potential effect modifier of GLP1-RA or SGLT2i treatment effect. Individuals with type 2 diabetes and history of cardiovascular disease showed a meaningful reduction in the risk of cardiovascular events, over an average follow-up time of three years of treatment with longer acting GLP1-RA or SGLT-2i drugs, which was not observed among patients with high cardiovascular risk and no prior cardiovascular events. Trends towards larger reductions in the risk of MACE were noted for SGLT2i drugs among patients with CKD compared to patients with mild or no impaired renal function, and for GLP1-RA and SGLT-2i drugs among patients with baseline HbA1c levels equal or greater than 8% compared to those with HbA1c levels less than 8%.

Our findings of a greater reduction in the risk of MACE with GLP1-RA or SGLT-2i treatment among patients with established cardiovascular disease are consistent with previous literature [7,8] and support recent position statements of major diabetes medical societies [28, 29]. However, the trials included in this meta-analysis enrolled mostly – and sometimes only – patients with established cardiovascular disease and used heterogeneous definitions of established cardiovascular disease across trials, which ultimately differed from the definition adopted in the clinical guidelines [30, 31-33]. Thus, even though our findings did not identify
a cardiovascular benefit among patients without established cardiovascular disease, additional evidence specifically targeting this population would be needed to conclude whether GLP1-RA or SGLT-2i drugs are effective or not for primary prevention of cardiovascular events. Furthermore, a low percentage of the participants enrolled in the included trials had characteristics similar to the patient populations with type 2 diabetes commonly treated in routine care [30, 34]. A recent study showed that if the selection criteria of the EXSCEL, SUSTAIN-6, and LEADER trials were applied to the real-world population, only 13-16% of patients with type 2 diabetes would have been eligible, with the exception of 43% for the REWIND trial [34]. Thus, additional research on the potential effects of GLP1-RA or SGLT-2i in real-world patients with or without history of established cardiovascular disease is warranted.

The risk reduction of major cardiovascular events in patients treated with SGLT-2i or GLP1-RA drugs was similar in those with or without history of heart failure. Because of lack of granularity in the published data, we could not explore if there was any relevant difference in risk of developing heart failure among patients with and without heart failure. The only trial that reported a subgroup analysis on the association between SGLT2i and hospitalization for heart failure or cardiovascular death, i.e., the DECLARE trial, suggested that the benefit could be similar among patients with and without history of heart failure [prior heart failure: 0.79 (0.63-0.99 vs. no history of heart failure 0.84 (0.72-0.99))] [28]. A recent trial confirmed that dapagliflozin led to a reduction of heart failure events and cardiovascular deaths in patients with chronic heart failure independently of type 2 diabetes [DAPA-HF: 0.75 (0.63 to 0.90)] [35, 36].

SGLT2i drugs appeared to reduce the cardiovascular risk by a larger magnitude in the group with CKD compared to the group with mild or no impaired renal function (18% vs 9%). A recent trial, targeting specifically patients with type 2 diabetes and chronic kidney disease, showed a risk reduction in MACE among patients treated with canagliflozin with estimates very close to our findings [CREDENCE: (HR 0.80; 95% CI, 0.67 to 0.95) vs our results: (HR 0.82; 95% CI, 0.69 to 0.97)] [37]. Meta-analyses of trials on GLP1 RA drugs stratified by CKD showed a moderate heterogeneity. In addition to an under-representation of patients with CKD in cardiovascular outcome trials assessing GLP1 RA, the source of this heterogeneity is likely to depend on a complex combination of factors that should be further investigated with individual-level data. Potential differential effects of GLP1 RA agents on renal outcomes due to structural differences should also be considered. [38, 39]. GLP1-RA and SGLT-2i drugs also showed trends towards larger reductions in the risk of MACE in patients with baseline HbA1c levels equal or greater than 8% compared to those with HbA1c levels less than 8% (SGLT-2i:16% vs 8%; GLP1-RA: 18% vs. 9%). We could not explore further the characteristics of patients with uncontrolled diabetes because of lack of information provided by clinical trials.

Our meta-analysis is the first to explore the potential effect modification of SGLT2i and GLP1 RA treatment with respect to MACE by multiple baseline characteristics, as identified by the cardiovascular outcome trials on these medication classes. Previous meta-analyses only focused on the assessment of effect modification by cardiovascular disease [7, 8, 40], or provided direct comparisons between drug classes primarily on MACE or heart failure [41, 42], without considering potential treatment effect heterogeneity. Our study has several limitations. First, in the context of this exploratory analysis, we did not adjust our test statistics to account for multiple comparisons. Thus, the potential GLP1-RA or SGLT2i treatment effect heterogeneity identified across groups of patients with and without established cardiovascular disease would not be deemed significant in the setting of adjusted p-values. However, our results are in line with current literature supporting the presence of effect modification by history of established cardiovascular disease [7,8]. Second, we could not explore the modification of the treatment effect analyzing more than one factor at the time due to the small number of trials available. Third, there was some heterogeneity in the definition and measurement of the effect modifiers across the included trials, which we could not account for in our analyses. Fourth, some extent of heterogeneity in the treatment effects of individual agents within the same class (especially GLP1-RA drugs) cannot be ruled out and could not be investigated within the current study. Finally, some of our subgroup meta-analyses might be still underpowered to detect treatment heterogeneity.

**Conclusions**

The overall benefits of SGLT2i or GLP1-RA drugs on major adverse cardiovascular events range between 11% and 12%. Among several clinically relevant baseline patients’ characteristics, history of established atherosclerotic cardiovascular disease appears to be the only modifier of the treatment effect of SGLT2i or GLP1-RA drugs with respect to major cardiovascular events, though more information on the effect of these agents is needed among patients without history of cardiovascular disease. A direction
toward larger benefits was observed among patients with baseline CKD for the SGLT-2i treatment, and among patients with baseline uncontrolled diabetes for both SGLT-2i or GLP1-RA drugs.

**Declarations**

**Ethics approval and consent to participate**

Not applicable

**Consent for publication**

Not applicable

**Availability of data and materials**

All data generated or analyzed during this study are included in this published article and its supplementary information files.

**Competing interests**

The authors declare that they have no competing interests

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**Authors' Contributions**

All individuals listed as authors qualify for authorship according to the ICMJE guidelines. E.D., E.P., A.S.K. designed the research. E.D., E.J., S.H. conducted the research. E.D. and J.M.F. performed or assisted in performing the statistical analysis of the data. E.D. and E.P. wrote the manuscript draft. E.D., E.P, and A.S.K. had primary responsibility for the final content. All authors contributed to the critical revision of the manuscript for important intellectual content and approved the final manuscript.

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Table 1. General characteristics of the 10 Randomized Control Trials included in the meta-analysis

| Trial name | Year | Drug class | Exp. Vs Control arms | Centers and Countries, n. | Primary endpoint and key secondary endpoint | Follow-up, median in years | Patients, n. Exp.: n. Control | Age, mean in years | Diabetes duration, median in years | Male, n (%) | BMI, kg/m² | HbA1c, median, % | Established CVD, n (%) | Established HF, n (%) | eGFR <60 ml/min per 1.73 m², n (%) |
|------------|------|------------|----------------------|----------------------------|-----------------------------------------------|--------------------------|-----------------------------|----------------|---------------------------------|-------------|-------------|-----------------|---------------------|-----------------|---------------------|
| ELIXA      | 2015 | GLP-1      | Lixisenatide vs placebo | 49 countries | 4-point MACE and expanded MACE | 2.1 | 3034:3034 | 60.3 | 9.3 | 4207 (69) | 30.2 | 7.6 | 6068 (100) | 1358 (22) | 1407 (23) |
| LEADER     | 2016 | GLP-1      | Liraglutide vs placebo | 410 sites in 32 countries | 3-point MACE and expanded MACE | 3.8 | 4668:4672 | 64.3 | 12.8 | 6003 (64) | 32.5 | 8.7 | 7598 (81) | 1305 (14) | 2158 (23) |
| SUSTAIN-6  | 2016 | GLP-1      | Semaglutide vs placebo | 230 sites in 20 countries | 3-point MACE and expanded MACE | 2.1 | 1648:1649 | 64.6 | 13.9 | 2002 (61) | 32.8 | 8.7 | 2735 (83) | 777 (24) | 939 (28) |
| EXCEL      | 2017 | GLP-1      | Exenatide vs placebo | 687 sites in 35 countries | 3-point MACE and MACE components | 3.2 | 7356:7386 | 62 | 12 | 5603 (62) | 31.8 | 8 | 10782 (73) | 464 (3) | 1129 (8) |
| HARMONY    | 2018 | GLP-1      | Albiglutide vs placebo | 610 sites in 28 countries | 3-point MACE and expanded MACE | 1.6 | 4731:4732 | 64.1 | 14.1 | 6569 (69) | 32.3 | 8.8 | 9463 (100) | 1922 (20) | 2222 (23) |
| REWIND     | 2019 | GLP-1      | Dulaglutide vs placebo | 371 sites in 24 countries | 3-point MACE and 7 secondary outcomes* | 5.4 | 4949:4952 | 66.2 | 9.5 | 5312 (54) | 32.3 | 7.3 | 3114 (31) | 853 (8.6) | 2199 (22) |
| PIONEER-6  | 2019 | GLP-1      | Semaglutide vs placebo | 214 sites in 21 countries | 3-point MACE and other CVD outcomes** | 2.6 | 1591:1592 | 66 | 14.9 | 2176 (68) | 32.3 | 8.2 | 2095 (85) | 388 (12.2) | 856 (27) |
| EMPA-REG OUTCOME | 2015 | SGLT-2 | Empagliflozin vs placebo | 580 sites in 42 countries | 3-point MACE and 4-point MACE | 3.1 | 4687:2333 | 63.1 | 10** | 3336 (71) | 30.7 | 8.1 | 7020 (100) | 706 (10) | 1819 (26) |
| CANVAS     | 2017 | SGLT-2    | Canagliflozin vs placebo | 667 sites in 30 countries | 3-point MACE and all-cause and CVD deaths | 2.4 | 5795:4347 | 63.3 | 13.5 | 6509 (64) | 32 | 8.2 | 6656 (66) | 1461 (14) | 2039 (20) |
| DECLARE    | 2019 | SGLT-2    | Dapagliflozin vs placebo | 882 sites in 33 countries | 3-point MACE and CVD mortality + HF hospitalizations | 4.2 | 8582:8578 | 63.9 | 11.0 | 10738 (63) | 32.1 | 8.3 | 6974 (41) | 1724 (10) | 1265 (7) |

Notes: DPP-4, Dipeptidyl peptidase-4 inhibitors; GLP-1, Glucagon-Like Peptide Receptor Agonists; Sodium-Glucose Cotransporter 2 Inhibitors; MACE, major adverse cardiovascular events; CVD, cardiovascular disease; HF, heart failure. 3-point MACE includes cardiovascular death, nonfatal myocardial infarction, and nonfatal stroke; 4-point MACE includes cardiovascular death, nonfatal myocardial infarction, nonfatal stroke and hospitalization for unstable angina.

* composite clinical microvascular outcome comprising diabetic retinopathy or renal disease; hospital admission for unstable angina; each component of the primary composite cardiovascular outcome; death; and heart failure requiring either hospital admission or an urgent visit requiring therapy.

** 57% of the randomized patients had duration of diabetes longer than 10 years.

*** secondary outcomes: expanded MACE (unstable angina resulting in hospitalization or heart failure resulting in hospitalization); a composite of death from any cause, nonfatal myocardial infarction, or nonfatal stroke; and the individual components of these composite outcomes.

# Age ≥50 yr and established CVD or chronic kidney disease.
Figure 1

Meta-analysis of the association between antidiabetic treatments and major adverse cardiovascular events (MACE) stratified by drug classes

| Trial      | Treatment (events, n) | Placebo (events, n) | HR (95% CI) | Weight |
|------------|-----------------------|---------------------|-------------|--------|
| GLP-1      | ELIXA 406/3034        | 399/3034            | 1.02 (0.89, 1.17) | 9.41   |
|            | LEADER 608/4668       | 694/4672            | 0.87 (0.78, 0.97) | 13.33  |
|            | SUSTAIN-6 108/1648    | 146/1649            | 0.74 (0.58, 0.95) | 3.35   |
|            | EXSCEL 839/7256       | 905/7396            | 0.91 (0.83, 1.00) | 16.56  |
|            | HARMONY 338/4731      | 428/4732            | 0.78 (0.68, 0.90) | 9.05   |
|            | REWIND 557/4653       | 632/4682            | 0.88 (0.79, 0.99) | 12.67  |
|            | PIONEER-6 61/1591     | 78/1592             | 0.79 (0.57, 1.10) | 1.89   |
| Subtotal   | (I-squared = 40.9%, p = 0.118) |            | 0.88 (0.82, 0.94) | 66.26  |
| SGLT-2     | EMPA-REG OUTCOME 282/2333 | 282/2333  | 0.66 (0.74, 0.99) | 8.51   |
|            | CANVAS 535/5795       | 476/4347            | 0.86 (0.76, 0.98) | 10.57  |
|            | DECLARE 756/8582      | 803/8578            | 0.93 (0.84, 1.03) | 14.66  |
| Subtotal   | (I-squared = 0.0%, p = 0.548) |            | 0.89 (0.83, 0.96) | 33.74  |
| Overall    | (I-squared = 21.2%, p = 0.248) |            | 0.88 (0.84, 0.92) | 100.00 |

NOTE: Weights are from random effects analysis

Patients with established cardiovascular disease

Patients at risk of cardiovascular events.
Figure 2

Subgroup meta-analysis of the association between antidiabetic treatments and MACE stratified by drug classes in patients with established cardiovascular disease and at risk of cardiovascular events.

Figure 3

Subgroup meta-analysis of the association between antidiabetic treatments and MACE stratified by drug classes in patients with normal or mild impaired kidney function.

Figure 4

Subgroup meta-analysis of the association between antidiabetic treatments and MACE stratified by drug classes in patients with uncontrolled diabetes (HbA1c>8%) and better controlled diabetes (HbA1c≤8%).

Supplementary Files

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