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Racial/ethnic differences in health insurance adequacy and consistency among children: Evidence from the 2011/12 National Survey of Children’s Health

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Abstract

Background: Surveillance of disparities in healthcare insurance, services and quality of care among children are critical for properly serving the medical/healthcare needs of underserved populations. The purpose of this study was to assess racial/ethnic differences in children’s (0 to 17 years old) health insurance adequacy and consistency (child has insurance coverage for the last 12 months).

Design and methods: We used data from the 2011/2012 National Survey of Children’s Health (n=79,474). Descriptive statistics and logistic regression analyses were conducted to examine the distribution and influence of several sociodemographic/family related factors on insurance adequacy and consistency across different racial/ethnic groups.

Results: Stratified analyses by race/ethnicity revealed that white and black children living in households at or below 299% of the Federal Poverty Level (FPL) were approximately 29 to 42% less likely to have adequate insurance compared to children living in families of higher income levels. Regardless of race/ethnicity, we found that children with public health insurance were more likely to have adequate insurance than their privately insured counterparts, while adolescents were at greater risk of inadequate coverage. Hispanic and black children were more likely to lack consistent insurance coverage.

Conclusions: This study provides evidence that racial/ethnic differences in adequate and consistent health insurance exists with both white and minority children being affected adversely by poverty. Establishing outreach programs for low income families, and cross-cultural education for healthcare providers may help increase health insurance adequacy and consistency within certain underserved populations.

Introduction

Recent estimates indicate that 3.3 million children under 18 years old are uninsured in the U.S.1 The lack of health insurance has deleterious effects on children’s health and well-being.2,3 Children who lack health insurance coverage are more likely to experience adverse health outcomes and have limited access to preventive services compared to those who are insured.3 Children without consistent health insurance (e.g., those with periods of no insurance) also have a greater likelihood for unmet health care needs and receipt of preventive services.4,6

Minority children, specifically, tend to have higher rates of suboptimal dental and medical health as well as no usual source of care when compared to other racial/ethnic groups of children.7 As it is evidenced by a recent U.S. Census Bureau report, 53% of uninsured children are Hispanic and black.8 Among those insured, Hispanic and older children are more likely to be underinsured.9 Recent studies indicate that black and Hispanic children have a lower likelihood of not receiving all needed medical/dental care and specialty care compared to white children.10 Efforts to reduce disparities in access to quality healthcare are reflected with the implementation of the Affordable Care Act (ACA) which is designed to promote access to equitable and more efficient health care. The ACA invested in initiatives that aim to increase the racial and ethnic diversity of health care providers and provide them with strengthened cultural competency training.11 However, despite these efforts, research studies show a wide gap in health care access, quality, and outcomes.12 Limited research exists on factors that may influence under-insurance and inconsistencies in healthcare coverage, and overall access to quality healthcare among children. Although most racial/ethnic minority children are eligible for Medicaid and Children’s Health Insurance Program (CHIP), many are not aware of their eligibility, while others may face barriers in enrolling in such programs due to limited English proficiency.13

The present study was designed to address this gap in research and provide population-based evidence for policy makers and potential interventions targeting children who are at risk for not having adequate or consistent health insurance.

The purpose of the present study was to examine racial/ethnic differences in adequacy and consistency of public or private health insurance, and assess associations of several demographic and family related factors with insurance adequacy and consistency across different racial/ethnic groups. We used data from the 2011/2012 National Survey of Children’s Health (NSCH).14
Materials and Methods

The 2011/2012 NSCH is a cross-sectional telephone survey of US households with at least one resident child aged 0 to 17 years old at the time of the interview. The goal of this survey is to assess the physical and emotional health of children, as well as factors that may relate to child well-being, including medical homes, family interactions, parental health, school and after-school experiences, and neighborhood characteristics. The NSCH uses the same sampling frame as the Centers for Disease Control and Prevention (CDC) National Immunization Survey (NIS). Using a complex survey design, the 2011/2012 NSCH surveyed the parents of a total of 95,677 children, or approximately 2000 per state. Details on the sampling design and procedures of this survey are presented at the 2011/2012 NSCH Guide (NSCH, 2011-2012).15

he analytic sample of this study (n=79,474) excluded uninsured children and missing data on variables for which it ranged from 0.05% to 9%. There were 3464 (3.6%) uninsured children who were excluded from this analytic sample since this study aimed to analyze the adequacy and consistency of health insurance coverage. This sample selection approach is in line with previous research in the field.9,17 Further, multicultural racial/ethnic groups (n=10,446) were excluded from the analytical sample since there were several groups which were consisted of small sample sizes. Responses of don't know and refuse to answer were considered to be missing data (n=2293).

Measures

The outcome variables in this study included adequacy of health insurance and consistency of health insurance. We included multiple dimensions in the definition of insurance adequacy and consistency as cited by the American Academy of Pediatrics and previous studies.9,16

The outcome variables of adequacy and consistency were created and provided for analysis in this study by the NSCH research team (NSCH, 2011-2012). Appendix A provides a summary of these outcome variables. The adequacy variable was created from parents' or guardians' responses to four questions: Does the child’s health insurance offer benefits or cover services that meet his or her needs? Does the child’s health insurance allow him or her to see the health care providers he or she needs? Not including health insurance premiums or costs that are covered by insurance, do you pay any money for the child’s health care? A positive answer to the third question led to a follow-up question, How often are these costs reasonable? If a parent or guardian indicated that the out-of-pocket costs were always or usually, reasonable and responded as always or usually to the other questions listed above, then the child was considered to have adequate insurance coverage. Other research studies have used the same approach in measuring adequacy.9,12 The combination of these survey questions led to an outcome variable that consisted of two categories: adequate (current insurance is adequate for child’s needs, n=61,268), and not adequate or underinsured (n=17,913). The outcome variable adequacy did not include uninsured children, refused to answer and missing records. The second outcome variable, consistency of health insurance, measured whether a child had consistent health insurance coverage during the past 12 months (NSCH, 2011-2012, www.childhealthdata.org). It was created by the NSCH team using the following survey questions: Does the child have any kind of health care coverage, including health insurance, prepaid plans such as HMOs, or government plans such as Medicaid? which consisted of three parts: 1) If YES, Is that coverage for him/her insured by Medicaid or the Children’s Health Insurance Program, CHIP?, 2) If YES, During the past 12 months, was there any time when he/she was not covered by any health insurance?, 3) If NO, During the past 12 months, was there any time when he/she had health care coverage? The combinations of these survey questions resulted in the dichotomous variable (yes/no) for consistency of health insurance in the past 12 months: consistent (child had consistent health insurance coverage during the past 12 months, n=75,509), and not consistent (n=3541) (NSCH, 2011-2012). The independent variables considered in this study were: race/ethnicity with three groups: white, black, Hispanic children; children’s age classified as 0-5, 6-11, and 12-17 years old; gender (male and female); type of health insurance included public insurance and private insurance; poverty level which was constructed based on the family’s income and included 0-199% Federal Poverty Level (FPL), 200-299% FPL, 300-399% FPL, 400% FPL; family structure consisted of two parent – biological, two parent – step family, single mother – no father, and other family type; and maternal health status (both mental and physical status). These independent variables were selected based on previous research in the field.18-22

Data analysis

Descriptive statistics and Chi-square analysis were conducted to describe the sample and test for significant associations between the two main outcome variables and sample characteristics. Stratified analysis was performed based on children’s racial/ethnic background to examine associations between adequacy and consistency of health insurance with children’s characteristics. Several logistic regression models were built separately for each racial/ethnic group to determine how children’s characteristics and those of their families influence their healthcare insurance adequacy and consistency status. All analyses were weighted to account for the complex survey design of the 2011-2012 NSCH. The STATA software was used to analyze the data.23

Results

Among the children who were insured (94.5%), 57.4% of them were privately insured whereas 37.1% were publicly insured. The uninsured children (3.6%) were not part of this study, and it was excluded from the analytic sample. Overall, most of the parents reported that their children’s insurance was adequate (77.4%) and consistent (95.5%) during the past year. Racial/ethnic differences in sociodemographic/family characteristics are presented in Table 1. A significantly larger percent of Hispanic (39.9%) and black (23.2%) children had public insurance, while most white children were privately insured (74.9%). More than half of the Hispanic and black children lived in households at or below the 299% FPL and this was significantly higher than their white counterparts. White and Hispanic children had also some of the highest rates of living with mothers not being at good/excellent mental and physical health.

Adequacy of health insurance

Table 2 displays the distribution of children’s characteristics, family structure, type of health insurance and their mother’s health status by adequacy of healthcare insurance. A significantly larger percent of 12-17 year-old children (26.4%) as well as males (23.8%), children of Hispanic origin (24.3%) and those with private insurance (26.8%) lacked adequate health insurance coverage. Similarly, a large percent of children living in families between 200-399% of the FPL (approximately 53.2%) and those with mothers not having good/excellent physical and mental health (25.8%)
did not have adequate health insurance in the past year. The association between race/ethnicity and insurance adequacy was not statistically significant.

Stratified logistic regression analyses by race/ethnicity (Table 3) revealed that black female children were 1.36 (95% Confidence Interval (CI): 1.12, 1.67) times more likely to have adequate insurance than their male counterparts. No significant associations between gender and insurance adequacy were found for the other groups. Further, black children living in a household with a single mother were 1.43 (95% CI: 1.13, 1.81) times more likely to have insurance adequacy relative to those residing in a two-parent household (biological parents). White and black children living in households at or below 299% of the FPL were approximately 29 to 42% less likely to have adequate insurance compared to children living in families at or above 400% FPL. Poverty level and insurance adequacy were not significantly associated within the Hispanic group. Furthermore, Hispanic and white children living with mothers of poor mental and physical health were less likely to have adequate health insurance compared to children whose mothers’ health status was good/excellent. All children (regardless of race/ethnicity) between the ages of 12-17 years old had a decreased likelihood of having adequate health insurance compared to younger children, whereas children with public insurance were at least twice as likely to have insurance adequacy compared to children with private insurance.

### Consistency of health insurance

Associations between children’s characteristics and consistency of health insurance revealed that Hispanic (9.1%) and black (7.9%) children, those with public insurance (11.3%) and children living in households at or below the 299% FPL were significantly less likely to have consistent insurance in the past 12 months (Table 4). A significantly higher percent of children living in families other than two parent-biological (percent ranged from 6.7 to 8.5%) had gaps in their insurance. Further, children whose mothers were not in good/excellent mental and physical health were less likely to report consistent healthcare insurance.

Table 5 presents findings related to frequency distribution and logistic regression analysis for health insurance consistency. Regardless of racial/ethnic background, children with public insurance were less likely to report consistency of health insurance compared to those with private insurance. Higher poverty levels for white and Hispanic children significantly decreased their likelihood for consistent health insurance status compared to those living at 400% FPL and above. Hispanic children whose mothers had poor mental and physical health were 0.64 (95% CI: 0.45, 0.91) times less likely to have consistent health insurance in comparison to those whose mothers were in good health. No significant associations were observed between age, gender and family structure with health insurance consistency for any of the racial/ethnic groups in the logistic regression models.

### Discussion

This study assessed associations between children’s sociodemographic and family characteristics with adequacy and consistency of health insurance coverage within the context of race/ethnicity. It builds on previous research on this subject and adds to the understanding of the role of these characteristics within the context of race/ethnicity in influencing healthcare insurance coverage.9,18,19

We found that adolescents across all racial/ethnic groups were less likely to have adequate insurance compared to younger children. This has been observed in previous studies and it may be an indicator of needed services that are not covered normally by insurance (e.g., mental health care).9 Although most states have introduced some form of CHIP to cover children less than five years of age, it may not be sufficient for older children. This highlights the need to focus on older children to ensure adequate

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Table 1. Racial/ethnic associations with socio-demographic/family characteristics.

|                     | White Unweighted Count (Weighted %) | Hispanic | Black | P-value |
|---------------------|-------------------------------------|----------|-------|---------|
| **Age**             |                                     |          |       |         |
| 0-5 years old       | 17,704 (57.0)                      | 4,360 (29.1) | 2,610 (13.9) | 0.0000 |
| 6-11 years old      | 18,825 (58.5)                      | 3,887 (25.9) | 2,860 (15.6) |       |
| 12-17 years old     | 22,853 (63.0)                      | 3,336 (20.9) | 3,039 (16.1) |       |
| **Gender**          |                                     |          |       | 0.2457 |
| Male                | 30,612 (60.0)                      | 6,035 (25.3) | 4,371 (14.7) |       |
| Female              | 28,705 (58.1)                      | 5,542 (25.1) | 4,135 (15.8) |       |
| **Type of healthcare insurance** |          |          |       |         |
| Public insurance    | 12,437 (59.1)                      | 6,149 (39.9) | 4,618 (23.2) | 0.0000 |
| Private Insurance   | 46,452 (74.9)                      | 5,187 (15.2) | 3,739 (9.9)  |       |
| **Poverty level**   |                                     |          |       | 0.0000 |
| 0-199% FPL          | 14,145 (39.8)                      | 6,606 (39.0) | 4,489 (21.2) |       |
| 200-299% FPL        | 10,056 (66.3)                      | 1,510 (19.8) | 1,258 (13.9) |       |
| 300-399% FPL        | 9,750 (74.0)                       | 1,066 (14.2) | 880 (11.8)  |       |
| 400% FPL and above  | 23,431 (79.6)                      | 2,401 (12.2) | 1,882 (8.2)  |       |
| **Family Structure**|                                     |          |       | 0.0000 |
| Two parents – biological | 45,243 (66.3)                     | 7,477 (25.2) | 3,269 (8.5)  |       |
| Two parents – step family | 3,931 (61.6)                     | 824 (22.1)  | 684 (16.3)  |       |
| Single mother – no father | 6,585 (38.4)                     | 2,451 (27.9) | 3,389 (33.7) |       |
| Other family type   | 3,586 (53.0)                       | 742 (20.6)  | 1,107 (26.4) |       |
| **Maternal health status (mental and physical)** |          |          |       | 0.0000 |
| BOTH Good/Excellent | 37,187 (67.8)                      | 5,143 (19.4) | 3,831 (12.8) |       |
| BOTH Not Good/Excellent | 18,826 (49.6)                     | 5,692 (33.5) | 3,559 (16.5) |       |
Table 2. Distribution of children's sociodemographic/family characteristics by health insurance adequacy in 2011/12 NSCH.

| Adequacy of health insurance: Unweighted Count (Weighted %) | Adequate | Not Adequate | P-value |
|-------------------------------------------------------------|----------|--------------|---------|
| **Age**                                                     |          |              |         |
| 0-5 years old                                               | 20,200 (81.1) | 4,391 (18.9) | 0.0000  |
| 6-11 years old                                              | 19,387 (76.1) | 6,096 (23.9) |         |
| 12-17 years old                                             | 21,681 (73.6) | 7,426 (26.4) |         |
| **Gender**                                                  |          |              | 0.0381  |
| Male                                                        | 31,520 (76.2) | 9,348 (23.8) |         |
| Female                                                       | 29,695 (77.6) | 8,544 (22.4) |         |
| **Race/ethnicity**                                          |          |              | 0.1283  |
| White                                                       | 45,853 (77.1) | 13,367 (22.9) |         |
| Hispanic                                                    | 8,835 (75.8) | 2,662 (24.2) |         |
| Black                                                       | 6,580 (78.1) | 1,884 (21.9) |         |
| **Type of healthcare insurance**                            |          |              | 0.0000  |
| Public insurance                                            | 19,577 (82.5) | 3,605 (17.5) |         |
| Private insurance                                           | 41,173 (73.2) | 14,173 (26.8) |         |
| **Poverty Level**                                           |          |              | 0.0000  |
| 0-199% FPL                                                  | 19,981 (78.7) | 5,095 (21.3) |         |
| 200-299% FPL                                                | 9,339 (74.1) | 3,417 (25.9) |         |
| 300-399% FPL                                                | 8,604 (72.7) | 3,054 (27.3) |         |
| 400% FPL and above                                          | 23,344 (77.7) | 6,547 (22.3) |         |
| **Family structure**                                        |          |              | 0.0110  |
| Two parents – biological                                    | 42,909 (76.0) | 12,907 (24.0) |         |
| Two parents – step family                                   | 4,165 (78.0) | 1,254 (22.0) |         |
| Single mother – no father                                   | 9,726 (78.8) | 2,655 (21.2) |         |
| Other family type                                           | 4,177 (78.4) | 1,009 (21.6) |         |
| **Maternal health status (mental and physical)**            |          |              | 0.0000  |
| BOTH Good/Excellent                                         | 36,445 (78.7) | 9,581 (21.3) |         |
| BOTH Not Good/Excellent                                     | 20,674 (74.2) | 7,285 (25.8) |         |

Table 3. Weighted analysis of health insurance adequacy and odds ratios (OR) (95% Confidence Intervals) by race/ethnicity among children in 2011/12 NSCH.

| Health Insurance Adequacy, White | Adequate | Not Adequate | OR (95% CI) | Adequate | Not Adequate | OR (95% CI) | Adequate | Not Adequate | OR (95% CI) |
|----------------------------------|----------|--------------|-------------|----------|--------------|-------------|----------|--------------|-------------|
| **Age**                          |          |              |             |          |              |             |          |              |             |
| Between 0-5                       | 81.5     | 18.4         | 1.00        | 82.0     | 18.0         | 1.00        | 79.8     | 20.2         | 1.00        |
| 6-11 years                       | 76.1     | 23.9         | 0.74 (0.66, 0.82) | 79.1     | 20.2         | 0.77 (0.61, 0.99) | 74.4     | 25.6         | 0.82 (0.63, 1.06) |
| 12-17 years                      | 74.1     | 25.9         | 0.69 (0.61, 0.76) | 74.1     | 23.9         | 0.69 (0.54, 0.91) | 71.9     | 28.1         | 0.65 (0.51, 0.84)  |
| **Gender**                       |          |              |             |          |              |             |          |              |             |
| Male                             | 77.0     | 23.0         | 1.00        | 76.0     | 24.0         | 1.00        | 74.3     | 25.7         | 1.00        |
| Female                           | 77.1     | 22.9         | 1.02 (0.94, 1.11) | 80.2     | 19.8         | 1.14 (0.93, 1.40) | 77.3     | 22.7         | 1.36 (1.12, 1.67)  |
| **Type of health insurance**     |          |              |             |          |              |             |          |              |             |
| Public insurance                 | 85.8     | 14.2         | 2.66 (2.27, 3.12) | 82.2     | 17.1         | 1.92 (1.44, 2.55) | 79.7     | 20.3         | 2.10 (1.62, 2.73)  |
| Private insurance                | 74.1     | 25.8         | 1.00        | 72.1     | 27.9         | 1.00        | 68.8     | 31.2         | 1.00        |
| **Poverty Level**                |          |              |             |          |              |             |          |              |             |
| 0-199% FPL                       | 79.8     | 20.2         | 0.68 (0.59, 0.79) | 78.8     | 21.2         | 0.93 (0.66, 1.31) | 77.4     | 22.6         | 0.58 (0.42, 0.81)  |
| 200-299% FPL                     | 73.1     | 26.9         | 0.64 (0.57, 0.73) | 76.6     | 23.4         | 1.10 (0.76, 1.58) | 75.7     | 24.3         | 0.71 (0.52, 0.98)  |
| 300-399% FPL                     | 72.9     | 27.1         | 0.72 (0.64, 0.81) | 77.2     | 22.8         | 0.86 (0.55, 1.34) | 67.8     | 32.2         | 0.83 (0.57, 1.20)  |
| 400% FPL and above               | 78.6     | 21.4         | 1.00        | 77.3     | 22.8         | 1.00        | 71.8     | 28.2         | 1.00        |
| **Family structure**             |          |              |             |          |              |             |          |              |             |
| Two parents – biological         | 76.5     | 23.5         | 1.00        | 74.2     | 25.8         | 1.00        | 73.4     | 24.6         | 1.00        |
| Two parents – step family        | 78.1     | 21.9         | 1.12 (0.95, 1.32) | 78.2     | 21.8         | 1.17 (0.80, 1.72) | 77.4     | 22.6         | 1.29 (0.89, 1.86)  |
| Single mother – no father        | 77.7     | 22.2         | 1.00 (0.87, 1.15) | 81.6     | 18.4         | 1.02 (0.79, 1.32) | 76.9     | 23.0         | 1.43 (1.13, 1.81)  |
| Other family type                | 80.0     | 19.9         | 0.88 (0.79, 0.97) | 77.7     | 22.3         | 0.98 (0.81, 1.18) | 73.5     | 26.5         | 1.21 (1.03, 1.42)  |
| **Maternal health status (mental and physical)**            |          |              |             |          |              |             |          |              |             |
| BOTH Good/Excellent              | 78.9     | 21.1         | 1.00        | 79.0     | 20.9         | 1.00        | 77.9     | 22.1         | 1.00        |
| BOTH Not Good/Excellent          | 73.2     | 26.8         | 0.68 (0.62, 0.74) | 77.3     | 22.8         | 0.73 (0.58, 0.91) | 74.2     | 25.8         | 0.85 (0.69, 1.06)  |
Table 4. Distribution of children's sociodemographic/family characteristics by health insurance consistency in 2011/12 NSCH.

|                        | Consistent | Inconsistent | P-value |
|------------------------|------------|--------------|---------|
| Age                    |            |              |         |
| 0-5 years old          | 23,290 (93.4) | 1,259 (6.6)  | 0.0802  |
| 6-11 years old         | 24,271 (93.6) | 1,181 (6.4)  |         |
| 12-17 years old        | 27,948 (94.5) | 1,101 (5.5)  |         |
| Gender                 |            |              | 0.8289  |
| Male                   | 38,963 (93.9) | 1,834 (6.1)  |         |
| Female                 | 36,476 (93.8) | 1,703 (6.2)  |         |
| Race/Ethnicity         |            |              | 0.0000  |
| White                  | 56,970 (95.5) | 2,172 (4.5)  |         |
| Hispanic               | 10,592 (90.9) | 876 (9.1)    |         |
| Black                  | 7,947 (92.1) | 493 (7.9)    |         |
| Type of Healthcare insurance |          |              | 0.0000  |
| Public insurance       | 20,894 (88.7) | 2,204 (11.3) |         |
| Private Insurance      | 54,016 (97.2) | 1,294 (2.8)  |         |
| Poverty Level          |            |              | 0.0000  |
| 0-199% FPL             | 22,879 (90.6) | 2,125 (9.4)  |         |
| 200-299% FPL           | 12,068 (93.0) | 659 (7.0)    |         |
| 300-399% FPL           | 11,334 (96.0) | 309 (4.0)    |         |
| 400% FPL and above     | 29,228 (98.3) | 448 (1.7)    |         |
| Family Structure       |            |              | 0.0000  |
| Two parents – biological | 53,828 (94.9) | 1,939 (5.1)  |         |
| Two parents – step family | 5,075 (92.0) | 336 (8.0)    |         |
| Single mother – no father | 11,449 (91.5) | 899 (8.5)    |         |
| Other family type      | 4,801 (93.3) | 341 (6.7)    |         |
| Maternal health status (mental and physical) | 0.0000 | 0.0000 | |
| BOTH Good/Excellent    | 44,509 (95.4) | 1,475 (4.6)  |         |
| BOTH Not Good/Excellent| 26,177 (91.8) | 1,723 (8.2)  |         |

Table 5. Weighted analysis of health insurance consistency and Odds Ratios (OR) (95% Confidence Intervals) by race/ethnicity among children in 2011-2012 NSCH.

|                        | Health Insurance Consistency, White | Health Insurance Consistency, Hispanic | Health Insurance Consistency, Black |
|------------------------|-------------------------------------|---------------------------------------|-------------------------------------|
|                        | Yes (%) | No (%) | OR (95% CI) | Yes (%) | No (%) | OR (95% CI) | Yes (%) | No (%) | OR (95% CI) |
| Age                    |         |        |            |         |        |            |         |        |            |
| 0-5 years old          | 91.5    | 8.4    | 1.00       | 87.8    | 12.2   | 1.00       | 84.4    | 15.5   | 1.00       |
| 6-11 years old         | 91.5    | 8.4    | 0.90 (0.70, 1.14) | 86.9    | 13.2   | 0.96 (0.68, 1.38) | 81.7    | 18.6   | 1.15 (0.71, 1.88) |
| 12-17 years old        | 92.2    | 7.7    | 1.03 (0.80, 1.32) | 87.9    | 12.1   | 0.95 (0.62, 1.46) | 79.6    | 20.4   | 0.89 (0.57, 1.39) |
| Gender                 |         |        |            |         |        |            |         |        |            |
| Male                   | 91.8    | 8.2    | 1.00       | 88.5    | 11.4   | 1.00       | 80.6    | 19.4   | 1.00       |
| Female                 | 91.9    | 8.0    | 0.98 (0.81, 1.20) | 86.5    | 13.4   | 1.10 (0.79, 1.52) | 83.8    | 16.2   | 0.76 (0.54, 1.07) |
| Type of health insurance |         |        |            |         |        |            |         |        |            |
| Public insurance       | 88.9    | 10.5   | 0.36 (0.27, 0.48) | 89.6    | 10.4   | 0.29 (0.18, 0.46) | 87.6    | 12.4   | 0.32 (0.19, 0.53) |
| Private insurance      | 97.5    | 2.48   | 1.00       | 95.9    | 4.1    | 1.00       | 96.4    | 3.6    | 1.00       |
| Poverty Level          |         |        |            |         |        |            |         |        |            |
| 0-199% FPL             | 84.8    | 15.2   | 0.32 (0.23, 0.45) | 85.5    | 14.5   | 0.48 (0.25, 0.93) | 78.8    | 21.2   | 0.79 (0.39, 1.58) |
| 200-299% FPL           | 89.9    | 10.0   | 0.31 (0.22, 0.43) | 88.0    | 11.9   | 0.26 (0.13, 0.56) | 80.6    | 19.4   | 0.64 (0.33, 1.25) |
| 300-399% FPL           | 94.3    | 5.6    | 0.45 (0.31, 0.65) | 91.1    | 8.9    | 0.56 (0.22, 1.46) | 89.7    | 10.3   | 0.50 (0.17, 1.49) |
| 400% FPL and above     | 97.4    | 2.6    | 1.00       | 93.1    | 6.8    | 1.00       | 96.3    | 3.7    | 1.00       |
| Family Structure       |         |        |            |         |        |            |         |        |            |
| Two parents – biological | 92.9    | 7.0    | 1.00       | 87.9    | 12.1   | 1.00       | 82.4    | 17.5   | 1.00       |
| Two parents – step family | 88.3    | 10.6   | 0.87 (0.61, 1.23) | 85.9    | 14.0   | 0.76 (0.43, 1.33) | 78.0    | 21.9   | 0.96 (0.52, 1.78) |
| Single mother – no father | 87.8    | 12.2   | 0.91 (0.70, 1.18) | 86.7    | 13.3   | 1.19 (0.83, 1.72) | 83.3    | 16.6   | 1.09 (0.69, 1.72) |
| Other family type      | 90.3    | 8.2    | 1.67 (0.47, 5.90) | 90.3    | 9.6    | 1.07 (0.12, 9.41) | 77.5    | 17.9   | 5.32 (0.59, 50.78) |
| Maternal health status (mental and physical) | 0.0000 | 0.0000 | |
| BOTH Good/Excellent    | 93.4    | 10.8   | 1.00       | 88.8    | 14.5   | 1.00       | 86.4    | 20.1   | 1.00       |
| BOTH Not Good/Excellent| 88.1    | 6.5    | 0.84 (0.69, 1.03) | 85.4    | 11.2   | 0.64 (0.45, 0.91) | 79.1    | 13.5   | 1.04 (0.73, 1.48) |
healthcare services. Furthermore, the findings of this study show that both black and white children living in low-income families may face challenges in obtaining adequate health insurance coverage. At a national level, over 65% of black children and around 30% of white children live in low-income households,24 with black children experiencing higher levels of financial burdens,18 which could contribute to both underinsurance and inconsistent coverage.9 Although previous studies consistently show that poverty negatively affects adequate health care coverage in black children,18,25 research also needs to address mechanisms by which poverty affects access to care and services among white children residing in low-income households.

In the present study, for Hispanic and white children poor maternal health was associated with a decreased likelihood of adequate health insurance. Generally, poor parental health and especially that of the mother has been linked to poor children’s health outcomes.26 In this study, access to care was the most significant factor for Hispanic and white children. The main reason for insurance inconsistency was expiration of benefits,18 which could contribute to both underinsurance and inconsistent coverage.18 Our findings on publicly insured children indicate the urgent need for improved focus on children’s continuation of healthcare services.

Regardless of race/ethnicity, we found that children with public healthcare insurance were more likely to have adequate insurance than their privately insured counterparts. This finding is consistent with previous research that has documented an increase in adequate insurance levels among publicly insured children.27,28 This could be due to Medicaid and CHIP Reauthorization Act of 2009 (CHIPRA). However, we found that children with public insurance were less likely to have consistent health insurance coverage compared to children with private health insurance. This highlights the fact that children on public health insurance may be more vulnerable to inconsistent health insurance coverage and hence intermittent access to healthcare. A recent cross-sectional study analyzed the reasons for lacking health insurance among minority children who were eligible for public insurance (Medicaid/CHIP) and found that the most common reason for insurance inconsistency was expiration of benefits and never reenrolled.18 Our findings on publicly insured children indicate the urgent need for improved focus on children’s continuation of healthcare services.

Although we did not find a significant association between race/ethnicity and insurance adequacy in bivariate analyses, Hispanic and Black children were more likely to lack consistent insurance coverage. These findings are in agreement with previous research studies that show minority children experiencing higher levels of inconsistent insurance coverage.9,18 An explanation of these findings can be drawn from the sociodemographic and family related characteristics of these children. We found that both black and Hispanic children had a high likelihood of living in poverty, having public insurance and residing in households with poor parental health, all of which can place a child at risk for inconsistent insurance coverage.18 In stratified analyses by race/ethnicity, we found that Hispanic and white children living in lower-income households were more likely to lack consistent insurance coverage. Studies have shown that although children of lower-income levels may be able to obtain health insurance via different programs (e.g., Children’s Health Insurance Program—CHIP), they are significantly more likely to be inconsistently insured.29 In previous studies, lack of consistent insurance among minority children has been associated with limited English proficiency, complex enrollment processes and families not being aware of their eligibility for CHIP programs.30 Further, a recent study indicated that expanding CHIP public insurance to minority children would increase coverage by 24.5%, improve access to care, and reduce disparities substantially.31 This could also be extended to white children residing in low-income households.

Several study limitations are present in this study. The NSCH is a cross-sectional survey and consequently cannot measure racial/ethnic differences in insurance adequacy and consistency longitudinally. The definitions of insurance adequacy and consistency capture access to healthcare partly as insurance coverage is multidimensional. Further, several measures used in this study were self-reported and were associated with telephone surveys which are subject to biases. For example, parental or guardian responses to some of the survey questions might be biased depending on their knowledge and perceptions of whether their children’s health insurance covers needed services. The strengths of this study include the use of a large nationally representative sample and measures of both insurance adequacy and consistency.

Conclusions

In this nationally representative sample of children, differences and similarities were observed in contributing factors to insurance adequacy and consistency among children of different racial/ethnic backgrounds. Low-income and poor maternal health adversely affected insurance consistency and adequacy for both minority and white children. Considering the lack of health insurance adequacy among adolescents, policy makers should consider extending health insurance coverage for this older group of children. Research studies on inequalities of healthcare coverage should also monitor under-served white populations of children as their insurance coverage is affected by similar factors as those for minority children. Elimination of inequalities may require targeted interventions that include cross-cultural education of healthcare providers, policy changes to grant low-income children with appropriate and reliable health insurance, research that may shed some light on the underlying mechanisms of underinsurance/inconsistency of insurance in minority populations, and an ongoing monitoring of disparities by health plans (e.g., CHIP, Medicaid, private insurance).

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Key words: Health insurance adequacy; health insurance consistency; race/ethnicity; children; sociodemographic and family factors.

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