ON A RANDOM GRAPH RELATED TO QUANTUM THEORY

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Abstract. We show that a random graph studied by Ioffe and Levit is an example of an inhomogeneous random graph of the type studied by Bollobás, Janson and Riordan, which enables us to give a new, simple, proof of their result on a phase transition.

1. Introduction

Ioffe and Levit [4] have recently introduced and studied a new random graph model called a quantum random graph, more precisely a quantum version of the classical Erdős–Rényi random graph, because of connections with a quantum model. (A random-cluster type modification of the graph studied in [4] and here is a Fortuin–Kasteleyn representation of a mean-field Curie–Weiss model in transverse magnetic field, see [1] and [4]. The physical interpretations of the parameters $\beta$ and $\lambda$ below then are inverse temperature and magnetic field strength.)

The purpose of this paper is to show that this random graph can be regarded as an instance of the general inhomogeneous model studied by Bollobás, Janson and Riordan [3], and that the general results on a phase transition in [3] enable us to give a new, simple, proof of the main result by Ioffe and Levit [4].

The random graph model defined by Ioffe and Levit [4] has three parameters, $\beta \in (0, \infty)$, $\lambda \in [0, \infty)$ and $n \in \mathbb{N}$, and we will denote the resulting random graph by $Q_n^{\beta,\lambda}$. (Ioffe and Levit [4] consider also the case $\beta = \infty$, which gives an infinite random graph, but we will not do so.) We will keep $\beta$ and $\lambda$ fixed and let $n \to \infty$. The parameter $n$ then measures the size of the graph; more precisely, as will be seen below, the number of vertices is random but roughly proportional to $n$; note, however, that typically $n$ is not exactly the number of vertices.

To define $Q_n^{\beta,\lambda}$, following [4], let $S_\beta$ be a circle of length $\beta$, i.e., the interval $[0, \beta]$ with the endpoints identified, and consider $n$ disjoint copies $S_1^\beta, \ldots, S_n^\beta$ of $S_\beta$. Next, let there be a Poisson process of “holes” on each circle $S_i^\beta$; the Poisson processes have intensities $\lambda$ and are independent. Let $M_i$ be the number of holes on $S_i^\beta$, thus $M_i \sim \text{Po}(\lambda \beta)$. The holes split each circle $S_i^\beta$...
into open intervals $I^k_i$, $k = 1, \ldots, M^*_i$, where $M^*_i := \max\{M_i, 1\}$, and we let these intervals be the vertices of $Q_{\beta, \lambda}^n$. (If $M_i = 0$, so there are no holes on $S^{\lambda}_i$, we regard the entire circle as an interval.) In other words, the vertex set of $Q_{\beta, \lambda}^n$ is $\{I^k_i : 1 \leq i \leq n, 1 \leq k \leq M^*_i\}$; the number of vertices in $Q_{\beta, \lambda}^n$ is thus $\sum_{i=1}^n M^*_i$.

Finally, for each unordered pair $(i, j)$, we consider a Poisson process $L_{ij}$ on $S^{\lambda}_i$, with intensity $1/n$ and independent of everything else; for each point $x$ in $L_{ij}$ we consider the corresponding points in $S^{\lambda}_i$ and take the two intervals $I^k_i$ and $I^{l_j}$ containing these points; we add an edge between these two intervals (regarded as vertices in $Q_{\beta, \lambda}^n$). Note that this yields a multigraph; there may be multiple edges. For the purpose of this paper, all multiple edges may be ignored; more precisely, we may merge each group of parallel edges into a single edge. Moreover, the expected number of multiple edges is easily seen to be $O(1)$, more precisely, it is less than $\beta^2/4$, so they may be ignored for almost all other purposes too.

Note that when $\lambda = 0$, the vertices are the $n$ circles $S^{\lambda}_i$, and $Q_{\beta, 0}^n$ is the classical random graph $G(n, p)$ with $p = 1 - \exp(-\beta/n) \approx \beta/n$.

We state the main result by Ioffe and Levit [4] in Section 2. Our formulation differs somewhat from [4], but the main features describing the phase transition of $Q_{\beta, \lambda}^n$ are the same. We give our proof in Section 4 after showing in Section 3 that $Q_{\beta, \lambda}^n$ is an instance of the inhomogeneous random graph in [3].

We denote the numbers of vertices and edges of a graph $G$ by $v(G)$ and $e(G)$, respectively. Moreover, for a subgraph $G$ of $Q_{\beta, \lambda}^n$, the vertices are intervals in some $S^{\lambda}_i$, and it is of interest to study the total length of these intervals, i.e., $\sum_{I \in V(G)} |I|$; we denote this sum by $\ell(G)$ and call it the length $\ell(G)$ of $G$.

All unspecified limits are as $n \to \infty$.

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2. Results

We first observe that the number of vertices in $Q_{\beta, \lambda}^n$ is random and, by the construction above, given by $v(Q_{\beta, \lambda}^n) = \sum_{i=1}^n M^*_i$, while the length $\ell(Q_{\beta, \lambda}^n) = \beta n$ is deterministic. Moreover, the number of edges $e(Q_{\beta, \lambda}^n) \sim Po(\beta^2/2n)$ (if we include possible multiple edges). It follows immediately from the law of large numbers, cf. (3.3) and (3.5) below, that, as $n \to \infty$,

$$v(Q_{\beta, \lambda}^n)/n \overset{p}{\to} \mathbb{E}M^* = \lambda \beta + e^{-\lambda \beta},$$

$$\ell(Q_{\beta, \lambda}^n)/n = \beta,$$

$$e(Q_{\beta, \lambda}^n)/n \overset{p}{\to} \beta/2.$$
Since the expected number of multiple edges is \( O(1) \), (2.3) holds whether we count multiple edges or not.

Denote the components of \( Q_n^{\beta,\lambda} \) by \( C_1, C_2, \ldots \), in order of decreasing number of vertices (breaking ties by any rule), with \( C_j(G) = \emptyset \) if \( Q_n^{\beta,\lambda} \) has fewer than \( j \) components. The main result of Ioffe and Levit [4] is that there is a phase transition similar to the well-known phase transition in the random graph \( G(n,p) \) [2], [5]: for certain values of \((\beta,\lambda)\), all components are of order smaller than \( n \), while for other values of the parameters, there is with high probability (exactly) one giant component with a positive fraction of all vertices. More precisely, we have the following result, essentially due to Ioffe and Levit [4] but in our formulation and with some details added.

**Theorem 2.1.** Let

\[
F(\beta, \lambda) := \frac{2}{\lambda} (1 - e^{-\lambda \beta}) - \beta e^{-\lambda \beta}.
\]

If \( F(\beta, \lambda) \leq 1 \), let \( \pi := 0 \), while if \( F(\beta, \lambda) > 1 \), let \( \pi > 0 \) be the unique positive solution to

\[
\frac{2\lambda + \pi}{(\lambda + \pi)^2} (1 - e^{-(\lambda + \pi)\beta}) - \frac{\lambda \beta}{\lambda + \pi} e^{-(\lambda + \pi)\beta} = 1.
\]

Then

\[
v(C_1)/n \overset{P}{\to} \rho := \frac{\pi \lambda \beta}{\lambda + \pi} \left( 1 - e^{-(\lambda + \pi)\beta} \right) + e^{-\lambda \beta} \left( 1 - e^{-\pi \beta} \right),
\]

\[
\ell(C_1)/n \overset{P}{\to} \pi \beta,
\]

\[
e(C_1)/n \overset{P}{\to} \zeta := \pi (1 - \pi/2) \beta,
\]

while \( v(C_2)/n \overset{P}{\to} 0 \), \( \ell(C_2)/n \overset{P}{\to} 0 \), \( e(C_2)/n \overset{P}{\to} 0 \).

For \( \lambda = 0 \), we interpret (2.4) as the limiting value \( 2\beta - \beta = \beta \); the results then simplify to standard results for \( G(n,p) \).

Note that if \( F(\beta, \lambda) \leq 1 \), then \( \rho = \zeta = \pi = 0 \), and thus all components are small (\( o_p(n) \) vertices and edges), while if \( F(\beta, \lambda) > 1 \), then there is with high probability (exactly) one giant component with a positive fraction of all vertices.

The constant \( \pi \) is by (2.7) and (2.2) the asymptotic fraction of the total length of \( Q_n^{\beta,\lambda} \) that belongs to \( C_1 \). Hence, \( \pi \) equals the asymptotic probability that a given (or random) point on \( \bigcup_1^n S_i^\beta \) belongs to \( C_1 \). This leads to the following corollary, due to Ioffe and Levit [4]. (Our \( \pi \) is denoted \( p(\beta, \lambda) \) in [4]. Note that our formulation is somewhat weaker than theirs.)

**Corollary 2.2.** The asymptotic probability that two random points on \( \bigcup_1^n S_i^\beta \) (chosen independently and uniformly) belong to the same component equals \( \pi^2 \). The same holds conditioned on \( Q_n^{\beta,\lambda} \) too, with convergence in probability.

To choose a random point on \( \bigcup_1^n S_i^\beta \) means that we choose a vertex in \( Q_n^{\beta,\lambda} \) with probability proportional to its length. There is a similar result (with
limit $\rho^2$) if we choose two vertices at random with the uniform distribution, see [3, Theorem 3.14(i)].

Remark 2.3. In the proof below based on general results in [3], $\pi$ is calculated from the survival probability in a multi-type branching process. Because of the special structure in our case, $\pi$ can also be interpreted as the survival probability in a standard, single-type, branching process, with offspring distribution mixed Poisson $\text{Po}(\Gamma_\beta)$ where $\Gamma_\beta = \min(\Gamma, \beta)$ and $\Gamma \sim \Gamma(2, \lambda^{-1})$, see [4]. The multi-type branching process describes asymptotically the process of finding new vertices in a component by successive exploration of the neighbors of the ones already found; the type is here the length of the vertex. The single-type branching process ignores the lengths; this can be done because when we reach a new vertex, on a circle not seen before, its length has the distribution $\Gamma_\beta$, independent of what has happened before.

To consider this single-type process simplifies some of the calculations below; for example, it is elementary to see that this process is supercritical if and only if $F(\beta, \lambda) > 1$. We prefer, however, to use the setup in [3] below and leave it to the reader to explore modifications involving also the single-type process.

We can also obtain other results on $Q^\beta_n$ from [3], for example the asymptotic distribution of vertex degrees (a certain mixed Poisson distribution) and the typical and maximal distances between two vertices in the same component, where distances are measured in the graph theory sense of the minimal number of connecting edges.

3. THE GENERAL INHOMOGENEOUS MODEL

The general inhomogeneous random graph $G^\V(n, \kappa)$ is defined as follows; see [3, Section 2] for further details. We proceed in two steps, constructing first the vertices and then the edges. Note that $n$ is a parameter measuring the size of the graph, and we are primarily interested in asymptotics as $n \to \infty$. (In general, as in Section 1, $n$ is not exactly the number of vertices, see [3] below. Moreover, there is no need for the parameter $n$ to be integer valued in general, although it is so in our application.)

A generalized vertex space $\V$ is a triple $(\S, \mu, (\x_n)_{n \geq 1})$, such that the following holds.

(i) $\S$ is a separable metric space.

(ii) $\mu$ is a (positive) Borel measure on $\S$ with $0 < \mu(\S) < \infty$.

(iii) For each $n$, $\x_n$ is a random sequence $(x_1, x_2, \ldots, x_{N_n})$ of $N_n$ points of $\S$ (where $N_n$ may be deterministic or random).

(Formally, we should write $\x_n = (x_1^{(n)}, \ldots, x_{N_n}^{(n)})$, say, as we assume no relationship between the elements of $\x_n$ for different $n$, but we omit this extra index.)

Let $M(\S)$ be the space of all (positive) finite Borel measures on $\S$, and equip $M(\S)$ with the standard weak topology: $\nu_n \to \nu$ iff $\int f \, d\nu_n \to \int f \, d\nu$
for all bounded continuous functions \( f : \mathcal{S} \to \mathbb{R} \). Let
\[
\mu_n := \frac{1}{n} \sum_{i=1}^{N_n} \delta_{x_i}
\]
where \( \delta_x \) is the Dirac measure at \( x \in \mathcal{S} \); thus \( \mu_n \) is a random element of \( M(\mathcal{S}) \). We will further assume that
\[
(iv) \quad \mu_n \xrightarrow{p} \mu, \text{ as elements of } M(\mathcal{S}).
\]
Recall that a set \( A \subseteq \mathcal{S} \) is a \( \mu \)-continuity set if \( A \) is (Borel) measurable and \( \mu(\partial A) = 0 \), where \( \partial A \) is the boundary of \( A \). The convergence condition \( [iv] \) is equivalent to the condition that for every \( \mu \)-continuity set \( A \),
\[
\mu_n(A) := \frac{\# \{ i \leq N_n : x_i \in A \}}{n} \xrightarrow{p} \mu(A); \quad (3.1)
\]
see \[3, Lemma A.2\].

Given a generalized vertex space \( \mathcal{V} \), we further assume that \( \kappa \) is a symmetric non-negative (Borel) measurable function on \( \mathcal{S} \times \mathcal{S} \). We let the vertices of the random graph \( G^V(n, \kappa) \) be the integers \( 1, \ldots, N_n \). For the edges we have two different versions. (See \[3, Section 2\] for further, different but asymptotically equivalent, versions). In both versions we assume that \( x_n \) is given and consider each pair of vertices \( (i, j) \) with \( i < j \) separately. In the first version we add an edge between \( i \) and \( j \) with probability \( p_{ij} := \min\{\kappa(x_i, x_j)/n, 1\} \). In the second version, we instead create a multigraph by adding a \( \text{Po}(\kappa(x_i, x_j)/n) \) number of edges between \( i \) and \( j \). In both versions, this is done independently (given \( x_n \)) for all pairs \( (i, j) \).

In order to avoid pathologies, we finally assume
\[
(v) \quad \kappa \text{ is continuous a.e. on } \mathcal{S} \times \mathcal{S}; \\
(vi) \quad \kappa \in L^1(\mathcal{S} \times \mathcal{S}, \mu \times \mu); \\
(vii)
\[
\mathbb{E} e(G^V(n, \kappa)) \xrightarrow{n \to \infty} \frac{1}{2} \int_{\mathcal{S}^2} \kappa(x, y) \, d\mu(x) \, d\mu(y). \quad (3.2)
\]
As shown in \[3, Remark 8.2\], \( (vii) \) follows from the other assumptions if \( \kappa \) is bounded and \( \text{Var}(N_n) = o(n^2) \).

Note that the number of vertices \( v(G^V(n, \kappa)) \) is roughly proportional to \( n \); more precisely, by \[3,4\],
\[
\frac{v(G^V(n, \kappa))}{n} = \frac{N_n}{n} = \mu_n(\mathcal{S}) \xrightarrow{p} \mu(\mathcal{S}). \quad (3.3)
\]

3.1. The Ioffe and Levit random graph. We now show that the construction of \( Q_n^{\beta, \lambda} \) in Section 1 is an instance of the general construction just given. For the space \( \mathcal{S} \) we may choose the set of all open intervals in \( \mathbb{S}_\beta \). (We regard the full circle \( \mathbb{S}_\beta \) as an “interval”, and thus an element of \( \mathcal{S} \).

The first phase of the construction defines the vertices of \( Q_n^{\beta, \lambda} \), with each vertex corresponding to an interval \( I_i \in \mathcal{S} \). In the second phase, edges are added as in the second version above, with \( \kappa(I, J) = |I \cap J| \) for two intervals \( I, J \subseteq \mathbb{S}_\beta \).
The construction of edges a.s. does not distinguish between intervals of full length $\beta$ obtained by from a circle with a single hole, and the full circle obtained when there is no hole at all. We thus identify the different intervals of length $\beta$ and let the precise definition of $S$ be $S := S_1 \cup S_2$, where $S_1 := \{S_\beta\}$ is a singleton and $S_2$ is the set of all (open) intervals in $S_\beta$ of length less than $\beta$. It is convenient to use the identification

$$S_2 = \{(x, \ell) : x \in S_\beta, 0 < \ell < \beta\},$$

where $x$ is the left endpoint and $\ell$ is the length of the interval. We can define $S$ as a metric space by taking the product topology on $S_2$ and letting both $S_1$ and $S_2$ be open in $S$; we omit the standard construction of a corresponding metric.

Remark 3.1. It might be more natural to define $S$, as a topological and metric space, as the cone obtained as a quotient space of the product $S_\beta \times (0, \beta]$ by identifying all $(x, \beta)$ to a single point. (This gives a space that can be identified with an open disc of radius $\beta$.) For our purposes, the construction is not very sensitive to the choice of topology on $S$, and both these choices work well.

Clearly, (i) holds. Moreover, if $x_n$ is the collection of all intervals $I_i$ (in some order), then (iii) holds.

Consider the Poisson process of holes on a single circle $S_\beta$. This yields a (finite) random family of intervals $I_1, \ldots, I_{M^*}$, i.e., a random subset of $S$. (In other terminology, a point process on $S$.) We let $\mu$ be the intensity measure of this random set; in other words, for every Borel set $A \subseteq S$, $\mu(A)$ is the expected number of the created intervals that belong to $A$:

$$\mu(A) := \mathbb{E} \# \{i \leq M^* : I_i \in A\}.$$ 

In particular, since the number of intervals $M^* = \max\{M, 1\}$ with $M \sim \text{Po}(\lambda \beta)$,

$$\mu(S) = \mathbb{E} M^* = \mathbb{E} M + \mathbb{P}(M = 0) = \lambda \beta + e^{-\lambda \beta}.$$  \hspace{1cm} (3.5)

Clearly, (ii) holds.

Since the construction of holes and intervals proceeds independently in the $n$ circles $S_i^\beta$, (3.1) follows, for every Borel set $A \subseteq S$, immediately by the law of large numbers for sums of i.i.d. random variables. Hence, (iv) holds.

As already remarked, the second stage in the construction of $Q_{n, \lambda}^{\beta, \Lambda}$ is as required, with $\kappa(I, J) = |I \cap J|$ for $I, J \in \mathcal{S}$. Clearly, $\kappa$ is continuous and bounded on $\mathcal{S}$, and thus the remaining conditions (v), (vi) and (vii) hold, using [3, Remark 8.2] for the latter.

The random graph $Q_{n, \lambda}^{\beta, \Lambda}$ is thus an instance of the inhomogeneous random graph in [3], and we may use the results of [3].

Note that $\kappa(I, J) = 0$ for some pairs of intervals $I, J$, but $\kappa(I, S_\beta) > 0$ for every $I \in \mathcal{S}$. Since $\mu(S_\beta) > 0$, it follows that $\kappa$ is irreducible in the sense of [3, Definition 2.10].
To calculate the measure \( \mu \), note first that
\[
\mu(\{S_\beta\}) = P(M^* = 1) = P(M \leq 1) = (1 + \lambda \beta)e^{-\lambda \beta};
\]
this is the point mass at \( S \).

On \( S_2 \) we use the coordinates \((x, \ell)\) as in \[3.4\]. An interval \( I = (x, x + \ell) \), with \( 0 < \ell < \beta \), is obtained as a piece if and only if both \( x \) and \( x + \ell \) are points in the Poisson process \( \mathcal{H} \) of holes, and further no point of \( \mathcal{H} \) falls in \((x, x + \ell)\). The intensity of \( \mathcal{H} \) is \( \lambda \), and given that there is a point at \( x \), the gap \( L \) to the next point (if we unwind the circle \( S_\beta \) to a line) has an exponential distribution with distribution function \( P(L > \ell) = e^{-\lambda \ell} \) and thus density function \( \lambda e^{-\lambda \ell} \). Consequently, the restriction of \( \mu \) to \( S_2 \) has the density \( \lambda^2 e^{-\lambda \ell} \, dx \, d\ell \). (Another, but related, argument to see this uses, twice, the fact that the Palm process of a Poisson process equals the process itself.)

We thus have the formula
\[
\int_S f \, d\mu = f(S_\beta)(\lambda \beta + 1)e^{-\lambda \beta} + \int_0^\beta \int_0^\beta f(x, \ell) \lambda^2 e^{-\lambda \ell} \, dx \, d\ell. \tag{3.6}
\]
Note that taking \( f = 1 \) yields \( \int_S d\mu = e^{-\lambda \beta} + \lambda \beta \) in accordance with \[3.5\].

4. **Proof of Theorem 2.1**

Since \( Q_n^{\beta, \lambda} \) thus is an instance of the general inhomogeneous graph \( G^\nu(n, \kappa) \), the results in \[3\] connect the asymptotic behaviour of \( Q_n^{\beta, \lambda} \) to properties of the integral operator
\[
Tf(x) = \int_S \kappa(x, y)f(y) \, d\mu(y)
\]
on \( L^2(S, \mu) \). For example, one of the main results \[3\, Theorem 3.1(i)] says that there is a giant component in \( Q_n^{\beta, \lambda} \) if and only if \( \|T\|_{L^2(S, \mu)} > 1 \). (For simplicity, we write \( \| \|_X \) also for the operator norm of an operator on a Banach space \( X \).)

**Lemma 4.1.** The operator norm \( \|T\|_{L^2(S, \mu)} \) equals \( F(\beta, \lambda) \) defined in \[2.6\].

**Proof.** We exploit the symmetry of the construction under the group \( R := \{\tau_y\}_{y \in S_\beta} \) of translations (or rotations, depending on your point of view) of the circle \( S_\beta \), where we define \( \tau_y(x) := x + y \) (mod \( \beta \)) for \( x, y \in S_\beta \). Each \( \tau_y \) acts naturally on \( S \) (with \( S_\beta \subset S_1 \) as a fixed point) and on \( L^2(S, \mu) \), and it is obvious that \( T \) is invariant under \( R \), i.e. that \( T \tau_y = \tau_y T \), \( y \in S_\beta \).

Note that a function \( f(I) \) is \( R \)-invariant if and only it depends on the length \( |I| \) only. Let \( L^2_R(S, \mu) \) be the subspace of \( L^2(S, \mu) \) consisting of \( R \)-invariant functions, and note that the \( R \)-invariance of \( T \) implies that \( T : L^2_R(S, \mu) \to L^2_R(S, \mu) \). Since \( \kappa \) is bounded, \( T \) is a Hilbert–Schmidt operator and thus compact on \( L^2(S, \mu) \), and there exists a unique positive normalized eigenfunction \( \psi \) with eigenvalue \( \|T\|_{L^2(S, \mu)} \). Since every translation \( \tau_y(\psi) \) evidently is another such eigenfunction, \( \tau_y(\psi) = \psi \) so \( \psi \in L^2_R(S, \mu) \). (More
precisely, \(\tau_y(\psi) = \psi\) a.e. for every \(y\), and it follows easily that \(\psi = \psi'\) a.e., where \(\tau_y(\psi') = \psi\) everywhere for each \(y\). Consequently,

\[
\|T\|_{L^2(S,\mu)} = \|T\|_{L^2_R(S,\mu)}. \tag{4.1}
\]

Let \(\hat{S} := (0,\beta]\). If \(g\) is any function on \(\hat{S} = (0,\beta]\), we let \(\hat{g}\) denote the function \(I \mapsto g(|I|)\) on \(S\). Thus \(\hat{g}\) is \(R\)-invariant, and every \(R\)-invariant function on \(S\) equals \(\hat{g}\) for some \(g\). Further, let \(\hat{\mu}\) be the measure on \(\hat{S}\) induced by \(\mu\) and the mapping \(I \mapsto |I|\) of \(S\) onto \(\hat{S}\); in other words,

\[
\int_{\hat{S}} g \, d\hat{\mu} = \int_{S} \hat{g} \, d\mu
\tag{4.2}
\]

for any non-negative function \(g\) on \(\hat{S}\). Explicitly, by \(3.6\),

\[
\int_{\hat{S}} g \, d\hat{\mu} = \int_{S} g(|I|) \, d\mu(I) = g(\beta)(\lambda \beta + 1)e^{-\lambda \beta} + \int_{0}^{\beta} g(\ell)\beta \lambda^2 e^{-\lambda \ell} \, d\ell. \tag{4.3}
\]

Equivalently, \(d\hat{\mu} = (\lambda \beta + 1)e^{-\lambda \beta} \delta_\beta + \beta \lambda^2 e^{-\lambda \ell} \, d\ell\). Clearly, the mapping \(g \mapsto \hat{g}\) defines a natural isometry of \(L^2_R(S,\mu)\) and \(L^2(\hat{S},\hat{\mu})\).

If \(I, J \in S\), the average of \(\kappa(I,\tau_y(J)) = |I \cap \tau_y(J)|\) over all translations \(\tau_y\) is \(|I| |J| / \beta\). Hence, if \(g \in L^2(\hat{S},\hat{\mu})\),

\[
T\hat{g}(I) = \int_{\hat{S}} \kappa(I,J)g(|J|) \, d\mu(J) = \int_{S} \frac{|I|}{\beta} \frac{|J|}{|J|} g(|J|) \, d\mu(J) = \int_{S} \frac{|I|}{\beta} g(\ell) \, d\hat{\mu}(\ell).
\tag{4.4}
\]

Define the kernel \(\tilde{\kappa}\) on \(\hat{S}\) by \(\tilde{\kappa}(x,y) := xy / \beta\) and let \(\tilde{T}\) be the corresponding integral operator. We then have, by \(4.3\), \(T\hat{g}(I) = \tilde{T}g(|I|)\), and thus

\[
T\hat{g} = (\tilde{T}g)'. \tag{4.5}
\]

Hence, the restriction of \(T\) to \(L^2_R(S,\mu)\) is unitarily equivalent to \(\tilde{T}\) on \(L^2(\hat{S},\hat{\mu})\), and thus, using \(4.1\),

\[
\|T\|_{L^2(S,\mu)} = \|T\|_{L^2_R(S,\mu)} = \|\tilde{T}\|_{L^2(\hat{S},\hat{\mu})}.
\]

Finally, the kernel \(\tilde{\kappa}\) has rank 1, and \(\tilde{T}g = \beta^{-1}\langle g, \iota \rangle_\iota\), where \(\iota(x) = x\). Hence, by \(1.3\) and an elementary integration,

\[
\|\tilde{T}\|_{L^2(\hat{S},\hat{\mu})} = \beta^{-1}\|\iota\|_{L^2(\hat{S},\hat{\mu})}^2 = \beta^{-1} \int_{0}^{\beta} x^2 \, d\hat{\mu}(x) = \beta(\lambda \beta + 1)e^{-\lambda \beta} + \int_{0}^{\beta} \lambda^2 e^{-\lambda \ell} \, d\ell = F(\beta, \lambda).
\]

By Lemma \(1.1\) and \(3.3\) Theorem 3.1, \(v(C_1)/n \overset{p}{\to} \rho\) for some \(\rho \geq 0\), where \(\rho = 0\) if \(F(\beta, \lambda) \leq 1\) and \(\rho > 0\) if \(F(\beta, \lambda) > 1\). Moreover, by \(3.3\) Theorem 3.6, \(v(C_2)/n \overset{p}{\to} 0\); hence also \(\ell(C_2)/n \overset{p}{\to} 0\) because \(\ell(C_2) \leq \beta v(C_2)\), and \(e(C_2)/n \overset{p}{\to} 0\) by \(3.3\) Proposition 8.11.
To find the constant $\rho$, let, as in \[3\], $\rho_\kappa(x)$, for $x \in \mathcal{S}$, be the survival probability for a multi-type Galton–Watson branching process with type space $\mathcal{S}$, starting with a single individual of type $x$ and where an individual of type $y$ has offspring according to a Poisson process on $\mathcal{S}$ with intensity $\kappa(y,z)\,d\mu(z)$. Then, see \[3\,\text{Theorem 6.2}\], $\rho_\kappa(x)$ is the maximal solution of

$$
\rho_\kappa(x) = 1 - e^{-T_\kappa(x)} = 1 - \exp\left(-\int_\mathcal{S} \kappa(x,y)\rho_\kappa(y)\right);
$$

(4.6)

further, if $\|T\| \leq 1$ then $\rho_\kappa(x) = 0$, while otherwise $\rho_\kappa(x) > 0$ for every $x \in \mathcal{S}$, and there is no other non-zero solution. Then,

$$
\rho = \int_\mathcal{S} \rho_\kappa(x)\,d\mu(x).
$$

(4.7)

Furthermore, by \[3\,\text{Theorem 3.5}\], $e(C_1)/n \xrightarrow{p} \zeta$, where

$$
\zeta = \frac{1}{2} \int_{\mathcal{S}^2} \kappa(x,y)\left(\rho_\kappa(x) + \rho_\kappa(y) - \rho_\kappa(x)\rho_\kappa(y)\right)\,d\mu(x)\,d\mu(y).
$$

(4.8)

Similarly, as an immediate consequence of \[3\,\text{Theorem 9.10}\],

$$
\ell(C_1)/n \xrightarrow{p} \int_\mathcal{S} |I|\rho_\kappa(I)\,d\mu(I).
$$

(4.9)

It thus remains to find $\rho_\kappa(x)$, and to compute these integrals.

By symmetry, $\rho_\kappa(I)$ depends on $|I|$ only, and thus $\rho_\kappa = \tilde{\phi}$ for some function $\phi$ on $\hat{\mathcal{S}}$. Then \[4.10\] and \[4.15\] yield

$$
\tilde{\phi} = 1 - \exp(-T_\tilde{\phi}) = (1 - \exp(\hat{T}\phi))^V
$$

and thus

$$
\phi = 1 - \exp(-\hat{T}\phi);
$$

(4.10)

in other words, $\phi$ is the corresponding function for the kernel $\hat{\kappa}$ on $\hat{\mathcal{S}}$. Since $\hat{\kappa}$ has rank 1, $\hat{T}\phi = \pi\iota$ for some $\pi \geq 0$, where as above $\iota(x) = x$. Consequently, \[4.10\] shows that

$$
\phi(x) = 1 - e^{-\pi x}.
$$

(4.11)

Since $\hat{T}\phi(x) = \int_\mathcal{S} \frac{\partial}{\partial x}\phi(y)\,d\hat{\mu}(y)$, the relation $\hat{T}\phi = \pi\iota$, or $\hat{T}\phi(x) = \pi x$, can be written

$$
\pi = \int_\mathcal{S} \frac{\partial}{\partial x}\phi(y)\,d\hat{\mu}(y) = \beta^{-1} \int_\mathcal{S} y\left(1 - e^{-\pi y}\right)\,d\hat{\mu}(y).
$$

(4.12)

Evaluating this integral by \[4.13\], we find the equation \[2.5\]. The results quoted from \[3\,\text{Theorem 6.2}\] above imply that if $F(\beta,\lambda) \leq 1$, then $\pi = 0$, while if $F(\beta,\lambda) > 1$, then $\pi > 0$; moreover, there is no other positive solution of \[2.5\], since otherwise we would have another positive solution of \[4.10\].

Next, by \[4.7\] and \[4.12\],

$$
\rho = \int_\mathcal{S} \rho_\kappa\,d\mu = \int_\mathcal{S} \tilde{\phi}\,d\mu = \int_\hat{\mathcal{S}} \phi\,d\hat{\mu}.
$$
By (4.11), (4.3) and an elementary integration, we find (2.6). Further, by (4.9) and (4.12),
\[
\ell(C_1)/n \xrightarrow{p} \int_{\mathcal{S}} x\phi(x)\,d\hat{\mu}(x) = \beta \pi,
\]
which yields (2.7). Finally, (4.8) and (4.2) yield
\[
\zeta = \frac{1}{2} \int \int_{\mathcal{S}^2} \hat{\kappa}(x,y) (\phi(x) + \phi(y) - \phi(x)\phi(y)) \,d\hat{\mu}(x)\,d\hat{\mu}(y)
\]
\[
= \frac{1}{2\beta} \left( 2 \int_{\mathcal{S}} x\phi(x)\,d\hat{\mu}(x) \int_{\mathcal{S}} y\,d\hat{\mu}(y) - \left( \int_{\mathcal{S}} x\phi(x)\,d\hat{\mu}(x) \right)^2 \right),
\]
which by (4.12) and the elementary \( \int_{\mathcal{S}} x\,d\hat{\mu}(x) = \beta \) yields (2.8).

**Proof of Corollary 2.2.** The sought probability is, given \( Q_{n}^{\beta,\lambda} \),
\[
\mathbb{P}(\text{same component} \mid Q_{n}^{\beta,\lambda}) = \sum_{j \geq 1} \ell(C_j)^2 \left( \frac{1}{\beta n} \right)^2 = \sum_{j \geq 1} \left( \frac{\ell(C_j)}{\beta n} \right)^2.
\]

By (2.7), the term with \( j = 1 \) tends to \( \pi^2 \) in probability. Moreover,
\[
\sum_{j > 1} \left( \frac{\ell(C_j)}{\beta n} \right)^2 \leq \frac{\ell(C_2)}{\beta n} \sum_{j > 1} \frac{\ell(C_j)}{\beta n} \leq \frac{\ell(C_2)}{\beta n} \xrightarrow{p} 0.
\]

Hence, the conditional probability in (4.13) tends to \( \pi^2 \) in probability. The unconditional probability is the expectation of (4.13), which converges to \( \pi^2 \) by bounded convergence.

**Remark 4.2.** The proof shows that the largest component in \( Q_{n}^{\beta,\lambda} \) asymptotically behaves as for the random graph defined using the kernel \( \hat{\kappa} \) on the space \( \mathcal{S} \) with the measure \( \hat{\mu} \). Nevertheless, these two inhomogeneous random graphs differ in other respects, for example in the number of triangles, see [3, Section 17].

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