Identifying compartments in ecological networks based on energy channels

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Abstract

It has been confirmed in many food webs that the interactions between species are divided into "compartments," that is, subgroups of highly interacting taxa with few weak interactions between the subgroups. Many of the current methods for detecting compartments in food webs are borrowed from network theory, which do little to improve our understanding of the mechanisms underpinning them. Therefore, a method based on ecological context is needed. Here, we develop a new method for detecting compartments in food webs based on the reliance of each node on energy derived from basal resources (i.e., producers or decomposers). Additional Monte Carlo simulations were conducted to test the significance of the compartmentalization. Further, we applied a food web dynamics model to test whether the effects of permutation would be retained within a single compartment. The proposed method identified significant compartments in 23 of the 28 empirical food webs that were investigated. We further demonstrated that the effects of node removal were significantly higher within compartments than between compartments. Our methods and results emphasize the importance of energy channels in forming food web structures, which sheds light on the mechanisms of self-organization within food webs.

KEYWORDS
compartmentalization, energy channel, food web, removal effects

1 INTRODUCTION

The structure of food webs is believed to affect community stability or species persistence (Dunne, 2006; May, 1973); however, it is still a central challenge to understand how food web structure influences ecosystem functions (Thompson et al., 2012). Many food webs have been showed to consist of compartments, which are subgroups of taxa, with many strong interactions occurring within the subgroups and few weak interactions between the subgroups (Cohen, 1978; Girvan & Newman, 2002; Krause, Frank, Mason, Ulanowicz, & Taylor, 2003; Pimm & Lawton, 1980; Rezende, Albert, Fortuna, & Bascompte, 2009; Yodzis, 1982), although the mechanisms of the compartmentalization remain underexplored. The idea of compartmentalization can be traced back to Herbert Simon's (1962) parable of watchmakers, where he showed that partitioning assembly is more efficient and argued that complex systems in nature could benefit from their hierarchical structure. For example, food webs are comprised of many energy flows. The coupling of these flows has been shown to promote food web stability (Rooney, McCann, Gellner, & Moore, 2006; Teng & McCann, 2004). In subsequent years, ecologists have used a wide variety of methods and definitions to search for compartments in food webs (Allen & Starr, 1982; Allesina, Bodini, & Bondavalli, 2005; Borrett, 2009).
In network theory, several methods have been developed in recent decades to detect community (i.e., compartment) structure in complex networks (Guimerà, Sales-Pardo, & Amaral, 2007; Newman, 2004a; Thébault, 2013). The basic methodological idea is to find a division of the network that maximizes modularity, which is a numerical index of how good a particular division is (Newman, 2004a). Modularity measures the fraction of within-community links minus the expected value of the same quantity in a network with the same community divisions but random connections between the nodes (Newman & Girvan, 2004). Using the methods, compartments have been detected in different kinds of networks, such as social networks (Newman & Girvan, 2004), the Internet (Simonsen, Maslov, Sneppen, & Eriksen, 2003), and biochemical networks (Ravasz, Somera, Mongru, Oliva, & Barabási, 2002). Ecologists have tried to apply these methods to ecological networks (Krause et al., 2003; Leger, Daudin, & Vacher, 2015; Stouffer & Bascompte, 2011); however, food webs have at least two special characteristics: (1) The distribution of the link weights is largely uneven (O’Gorman, Jacob, Jonsson, & Emmerson, 2010); and (2) food web links have explicit directions from low to high trophic levels. These characteristics provide a cautionary note when interpreting results using methods designed for undirected and unweighted networks (Zhao et al., 2016). Moreover, the methods borrowed from network theory do little to improve our understanding of the underlying mechanisms of compartmentalization in food webs. Therefore, a method based on ecological context is urgently needed.

An energy channel is a collection of energy or material fluxes that start with a basal resource and end with a top predator (Moore & de Ruiter, 2012), which plays an important role in maintaining both food web structure and functioning. Different energy channels within the same food web may exhibit different traits, for example turnover rate (Rooney et al., 2006). For example, in marine food webs, phytoplankton energy channels were shown to have consistently higher turnover rates than detrital energy channels (Rooney, McCann, & Moore, 2008). The coupling of fast (e.g., phytoplankton) and slow (e.g., detrital) channels is believed to contribute to food web stability (Rooney & McCann, 2012; Rooney et al., 2006, 2008). Based on taxonomical aggregation, two compartments were defined: the detritus-based “brown” subweb and the producer-based “green” subweb (Butler, Gotelli, & Ellison, 2008; Zou, Thébault, Lacroix, & Barot, 2016). However, this separation is oversimplistic, and the compartmentalization based on energy channels is not fully explored. Here, we propose a new method for detecting compartments, where we consider one or several energy channels as a compartment. Our first task in this study was to test whether empirical food webs could be significantly compartmentalized using our method; that is, whether the modularity of the compartmentalization was significantly higher than what was likely to occur by chance.

Theoretically, compartmentalization is believed to greatly increase the stability and persistence of food webs (Krause et al., 2003; Pimm, 1979; Stouffer & Bascompte, 2011). One possible mechanism is that the impacts of a disturbance can be buffered by being retained within a single compartment and rarely spread to other compartments (Krause et al., 2003), although this hypothesis has not been tested well. In the study by Krause et al. (2003), an assumption was made to mimic the effects of species loss: After a taxon was removed, all interactions with this taxon were removed, and the predators on this taxon had their interaction strengths associated with this taxon redistributed proportionally to their interactions with other prey. This assumption is simple (as admitted by Krause et al.). It cannot track the biomass change after taxon removal, as well as the change in interaction strength associated with the biomass change. Moreover, they compared the removal effects of only two taxa, which may underrepresent the overall pattern. Our second task was to test the hypothesis that compartmentalization could increase food web stability using a dynamical approach. Finally, we compared our algorithm based on energy channels with some other commonly used compartment detection algorithms.

## 2. MATERIALS AND METHODS

### 2.1. Quantitative food webs

We analyzed 28 of the 50 aquatic food webs from a recently published database (see Table 1; Salas & Borrett, 2011; Borrett, 2013), which is archived in the "enR" package (Borrett & Lau, 2014) in R. Twenty food webs were excluded because they contained less than twenty taxa and thus were deemed too small. The other two food webs were excluded because they did not include microbial processes. The data for each food web include a list of taxa, the carbon biomass of each taxon (g C m⁻²), the carbon per unit time of import, export, and respiration of each taxon (g C m⁻² day⁻¹), and the carbon flux between a pair of taxa (g C m⁻² day⁻¹). Our focal food webs exhibit a wide range of taxon richness (5 = 23–124). Nodes represent species, trophic guilds, functional groups, or nonliving components of the system in which matter is stored. The ecosystems were considered to be in steady state, as in previous studies (Borrett, 2013; Borrett & Salas, 2010; Ulanowicz, 2004), which means that the flows of energy entering and leaving a given node are equal. Due to the errors and approximations associated with some measures, however, some food webs were initially unbalanced, that is, energy entering a taxon did not necessarily balance the output exactly. These food webs were balanced using the AVG2 algorithm using established procedures in Matlab 7.12.0 (Allesina & Bondavalli, 2003).

### 2.2. Measure of modularity

The classic definition of modularity proposed by Newman (2004a,b) was extended by Arenas, Duch, Fernández, and Gómez (2007) to the scenario of weighted directed networks as follows:

$$Q = \frac{1}{2W} \sum_{i} \sum_{j} \left( w_{ij} - \frac{w_{ij}^\text{in} w_{ij}^\text{out}}{W} \right) \delta (c_i, c_j)$$  

(1)
TABLE 1 Compartment analysis of 28 empirical food webs. For each food web, S is the number of taxa; $S_c$ is the number of detected compartments; Q is the empirical modularity; p is the probability that the modularity of a simulated food web in Monte Carlo simulations is higher than Q; and the last two columns indicate the mean and standard error (SEM) of modularity in Monte Carlo simulations. Significant Q values are marked in bold, where the significance level was $\alpha = 0.05/28 = 0.0018$ after Bonferroni correction.

| Index | Name                                | S   | $S_c$ | Q     | p        | Monte Carlo |
|-------|-------------------------------------|-----|-------|-------|----------|-------------|
|       |                                     |     |       |       |          | Mean        | SEM         |
| 1     | Chesapeake Bay                      | 36  | 3     | 0.39  | <.001    | -0.0052     | 0.0045      |
| 2     | Cypress (dry)                       | 68  | 3     | 0.13  | .006     | 0.0004      | 0.0005      |
| 3     | Cypress (wet)                       | 68  | 3     | 0.15  | .009     | -0.0002     | 0.0005      |
| 4     | Florida Bay (dry)                   | 124 | 4     | 0.10  | <.001    | 0.0081      | 0.0034      |
| 5     | Florida Bay (wet)                   | 124 | 4     | 0.10  | <.001    | 0.0008      | 0.0035      |
| 6     | Georges Bank                        | 31  | 2     | 0.10  | <.001    | -0.0020     | 0.0033      |
| 7     | Graminoids (dry)                    | 66  | 3     | 0.05  | <.001    | 0.0019      | 0.0023      |
| 8     | Graminoids (wet)                    | 66  | 3     | 0.03  | <.001    | 0.0008      | 0.0017      |
| 9     | Gulf of Maine                       | 31  | 2     | 0.10  | .001     | -0.0006     | 0.0033      |
| 10    | Lake Oneida (pre-ZM)                | 74  | 2     | 0.08  | <.001    | -0.0008     | 0.0004      |
| 11    | Lake Oneida (post-ZM)               | 76  | 3     | 0.08  | .018     | -0.0007     | 0.0020      |
| 12    | Mangroves (dry)                     | 94  | 3     | 0.02  | <.001    | -0.0009     | 0.0005      |
| 13    | Mangroves (wet)                     | 94  | 3     | 0.03  | <.001    | 0.0001      | 0.0005      |
| 14    | Middle Atlantic Bight               | 32  | 2     | 0.10  | .001     | -0.0013     | 0.0029      |
| 15    | Narragansett Bay                    | 32  | 2     | 0.10  | .019     | -0.0026     | 0.0025      |
| 16    | Neuse Estuary (early summer 1997)   | 24  | 3     | 0.27  | <.001    | 0.0100      | 0.0059      |
| 17    | Neuse Estuary (late summer 1997)    | 27  | 3     | 0.29  | <.001    | -0.0009     | 0.0069      |
| 18    | Neuse Estuary (early summer 1998)   | 24  | 3     | 0.33  | <.001    | -0.0063     | 0.0060      |
| 19    | Neuse Estuary (late summer 1998)    | 25  | 3     | 0.24  | <.001    | -0.0057     | 0.0063      |
| 20    | Northern Benguela Upwelling         | 23  | 2     | 0.08  | .004     | 0.0054      | 0.0044      |
| 21    | Southern New England Bight          | 33  | 2     | 0.11  | <.001    | -0.0028     | 0.0034      |
| 22    | St. Marks Seagrass, site 1 (January)| 43  | 4     | 0.20  | <.001    | 0.0018      | 0.0021      |
| 23    | St. Marks Seagrass, site 1 (February)| 45  | 4     | 0.16  | <.001    | -0.0003     | 0.0024      |
| 24    | St. Marks Seagrass, site 2 (January)| 40  | 3     | 0.19  | <.001    | -0.0004     | 0.0020      |
| 25    | St. Marks Seagrass, site 2 (February)| 46  | 3     | 0.17  | <.001    | -0.0015     | 0.0022      |
| 26    | St. Marks Seagrass, site 3 (January)| 32  | 4     | 0.21  | .001    | -0.0011     | 0.0024      |
| 27    | St. Marks Seagrass, site 4 (February)| 46  | 3     | 0.20  | <.001    | -0.0028     | 0.0033      |
| 28    | Sylt-Romo Bight                     | 59  | 5     | 0.20  | <.001    | 0.0026      | 0.0033      |

where $w_{ij}$ is the strength of the link (i.e., the carbon flux) from $i$ to $j$, and $w_{in}$ and $w_{out}$ are the input and output strengths of nodes $i$ and $j$, respectively:

$$w_{out} = \sum_j w_{ij}$$

(2)

$$w_{in} = \sum_i w_{ij}$$

(3)

The total strength of the link is:

$$2W = \sum_i w_{out} = \sum_i w_{in} = \sum_i \sum_j w_{ij}$$

(4)

The Kronecker delta function $\delta(c_i, c_j)$ takes the value of 1 if $c_i = c_j$ and otherwise 0, where $c_i$ is the community to which node $i$ is assigned.
Nonzero values of $Q$ indicate deviations from randomness, and a higher value of $Q$ indicates the food web is more compartmentalized.

2.3 | Compartment identification

Four steps were applied in our method to detect the compartments of a food web. First, a food web with $S$ taxa can be represented by an $S$-by-$S$ flow matrix $F = [F_{ij}]$; the value of the element $F_{ij}$ means the amount of carbon passing from taxon $i$ to taxon $j$ per unit area and time. Based on the flow matrix (Figure 1a), we calculated the reliance of each node on energy derived from producers or decomposers, that is, the proportion of carbon derived from basal nodes, $\%BR_r$ (Figure 1b), using a method proposed by Rooney et al. (2008):

$$\%BR_r = \sum_{i=1}^{n} P_{C} \times \%BR_r$$

where $n$ is the number of resources consumed by the consumer, $P_{C}$ is the proportion of the consumer diet accounted for by a resource, and $\%BR_r$ is the proportion of carbon derived from the basal resource in the resource being consumed.

Second, the nodes with the same dominant energy origins were put into the same subgroup. For example, in the Chesapeake Bay ecosystem, five subgroups were formed (Figure 1c). Here, each subgroup corresponds to one of the five basal nodes, and the nodes in a given subgroup obtained most of their energy from the same basal node.

Third, we tested whether the subgroups could be optimized by merging them into one compartment. Here, we calculated the difference in modularity $\Delta Q = Q_{after} - Q_{before}$, where $Q_{before}$ and $Q_{after}$ equate to the modularity before and after merging, respectively. By trying different combinations of subgroup pairs, we created a matrix, with the element $\Delta Q_{ij}$ indicating the difference in modularity after merging subgroup $i$ and subgroup $j$ (Figure 1d).

Fourth, if $\Delta Q > 0$, the merging was accepted as it could increase $Q$ and otherwise, the merging was rejected (Figure 1e). Repeat steps 3 and 4 until all values of $\Delta Q \leq 0$. With this four-step method, we were

![Figure 1](image-url)
able to detect the number of compartments and nodes within them for all the 28 food webs that we investigated.

2.4 Monte Carlo simulations

To test whether the concentration of interactions within identified compartments was greater than what was likely to occur by chance, we determined the statistical significance of the modularity using Monte Carlo simulations. For each simulation, we randomly reassigned interactions, constraining the sum of each column in the carbon-flow matrix to be equal to that of the original food web (where rows represented prey and columns represented predators), as in Krause et al. (2003). The detailed approach is as follows: (1) for column i (i.e., for predator i), we first calculated the sum of the interaction weight \( W_i \); (2) to determine the prey of predator \( i \) in the new food web, we chose \( n_i \) prey randomly from a pool of \( S-1 \) species where \( n_i \) is the number of prey for predator \( i \) in the original food web, and \( S-1 \) is the number of species excluding predator \( i \); (3) we generated \( n_i \) random numbers, \( x_1, x_2, ..., x_{n_i} \), each of which varied from 0 to 1; and (4) the weight of interactions between the \( j \)-th prey and predator \( i \) should be \( W_i x_j / \sum x_j \). This ensured that the simulated food webs had the same number of predators, the same number and weight of interactions associated with a predator, and the same number and weight of realized interactions as the original food web. We then calculated the modularity of the simulated food web on the compartments obtained from the method we proposed. This process was repeated 1,000 times to obtain a sampling distribution of modularity against which we compared the empirical modularity.

2.5 Food web dynamics

To test the hypothesis that compartmentalization would retain the impacts of a disturbance within a single compartment and with less spread to other compartments, we simulated the effects of randomly removing one node, by employing a food web dynamics model (Zhang, Zhao, Tian, & Huang, 2016; Zhao et al., 2016). This model was constructed based on energy budgets indexing the carbon fluxes entering and leaving each taxon. The taxa in each of the studied ecosystems can be divided into four categories: producers, consumers, decomposers, and detritus.

The change in biomass of producers can be described as:

\[
\frac{dB_i}{dt} = r_i B_i (1 - \sum_{j\in \text{pro}} B_j / K) - \sum_{j\in \text{pred}} \Phi_{ji} B_j - d_i B_i
\]

Here, "pro" are producer taxa, "herbi" are herbivorous taxa, \( r_i \) is the maximum specific or intrinsic growth rate, \( K \) is the carrying capacity, \( d \) is the specific death rate, and \( \Phi_{ji} \) is the functional response when taxon \( j \) consumes taxon \( i \), which was set to follow a nonlinear form as follows (Hudson & Reuman, 2013):

\[
\Phi_{ji} = \frac{y_{ij} B_j^3}{H_j^3 + q_j B_j H_j + \sum_{k\in \text{res}} \omega_{jk} B_k^2}
\]

Here, \( y_{ij} \) is the maximum consumption rate of taxon \( j \), \( \omega_{jk} \) is the preference of taxon \( j \) for taxon \( i \), \( H_i \) is the half-saturation density, \( q_j \) is the predator interference coefficient, and \( h \) is the hill exponent that regulates the shape of the curve from Holling type II (\( h = 1 \)) to Holling type III (\( h = 2 \)).

The change in biomass of consumers (including herbivores and predators) can be depicted as:

\[
\frac{dB_i}{dt} = \sum_{j\in \text{res}} a_{ij} B_j B_i - \sum_{j\in \text{pred}} \Phi_{ji} B_j - x_i B_i
\]

where, "res" are resource taxa, "pred" are predatory taxa, \( a \) is the assimilation efficiency, and \( x \) is the respiration rate.

The change in biomass of decomposers can be depicted as:

\[
\frac{dB_i}{dt} = \sum_{j\in \text{det}} a_{ij} B_j B_i - \sum_{j\in \text{pred}} \Phi_{ji} B_j - x_i B_i
\]

where "det" are detrital taxa.

In some food webs, detritus has been divided into separate taxa. The change in biomass of each detrital taxa can be described as:

\[
\frac{dB_i}{dt} = \sum_{j\in \text{pro}} a_{ij} B_j B_i + \sum_{j\in \text{con}} (p_{ij} e_{ij} B_j B_i) + \sum_{j\in \text{res}} c_{ij} B_j - \sum_{j\in \text{dec}} \Phi_{ji} B_j - \sum_{j\in \text{det}} c_{ij} B_i
\]

The setting and calculation of the parameters from carbon flux data can be found in Appendix S1 and Table S1. We used the Adaptive Runge–Kutta method to perform 1,000 numerical simulations for each food web. In each simulation, the empirical biomass data were employed to give the initial biomass values. First, 1,000 days were simulated to allow transient dynamics caused by initial effects to settle down (Hudson & Reuman, 2013). The average biomass density for each node during the next 1,000 simulated days was recorded as \( B_i^+ \).

We then randomly removed a node and simulated another 1,000 days, recording the average biomass density for each node during this period as \( B_i^- \).

2.6 Removal effects

The removal effect (RE) value of the loss of node \( k \) on node \( i \) was measured according to Brose, Berlow, and Martinez (2005) as:

\[
RE_i = \log \left( \frac{(B_i^+ + 1)(B_i^- + 1)}{B_i + 1} \right)
\]

where \( B_i \) is the biomass density of the node to be removed before its removal, while \( B_i^+ \) and \( B_i^- \) are the biomass densities of \( i \) with and without \( k \), respectively. Adding one to the biomass densities prevents very low \( B_i \) from generating artificially large values of RE (Brose...
et al., 2005). The within-compartment RE is the average RE value for all the nodes within the same compartment as k, and the between-compartment RE is the average RE value for all the nodes in different compartments to k.

2.7 | Comparison with other algorithms

To test the performance of our algorithm based on energy channels, we compared the modularity values that were generated with two other commonly used algorithms. The Girvan–Newman algorithm (Newman & Girvan, 2004) first estimates the edge betweenness score as the number of shortest paths through an edge. This algorithm then gradually removes the edge with the highest edge betweenness score to produce a dendrogram of removals. This dendrogram is then cut at the point which gives the highest value of modularity. The second algorithm finds densely connected subgraphs by performing random walks (Pons & Latapy, 2005), where each walk tends to stay inside communities instead of jumping to other communities.

2.8 | Statistical analyses

To compare the two categories of RE values (within-compartment, and between-compartment), we employed paired t test. To compare the performance of the three compartmentalization algorithms, we employed a linear mixed effects model (LME) with a maximum-likelihood estimator (function “lme” with “method = ML” within the “nlme” package in R). Food web identity was included in the model as a random factor to correct for differences between study systems. Post hoc comparisons were applied using the Tukey HSD test at α = 0.05 level of significance (function “glht” within the “multcomp” package in R).

3 | RESULTS

First, we showed a detailed example of one food web: the Chesapeake Bay. This ecosystem contained 31 consumer and decomposer taxa and five basal taxa: phytoplankton, benthic diatoms, dissolved organic nutrient, suspended particulate nutrient, and sediment particulate nutrient (Figure 2). Among all the 31 consumer and decomposer taxa, the energy of 13 taxa was mainly supplied by sediment particulate nutrient, while suspended particulate nutrient was the major supplier for another 12 taxa (Figure 1b), phytoplankton was the major supplier for four taxa, dissolved organic nutrient was the major supplier for two taxa, and benthic diatoms were not a major supplier for any taxa. Three compartments were detected, one of which was mainly sustained by suspended particulate nutrient and phytoplankton; another compartment was mainly sustained by sediment particulate nutrient but also contained benthic diatoms, while the final compartment contained just three taxa: free bacteria, heterotrophic microflagellates, and dissolved organic nutrient.

The compartments of all 28 food webs (Appendix S2) showed three types of compartmentalization. In the first type, pelagic taxa (supported by phytoplankton) and benthic taxa (supported by sediment bacteria or benthic algae) were separated. Sometimes, the taxa supported by sediment bacteria were also separated from the taxa supported by benthic algae. Fifteen ecosystems showed this type of compartmentalization, including Chesapeake Bay, Georges Bank, Gulf of Maine, Lake Oneida (pre-ZM), Lake Oneida (post-ZM), Mangroves (dry), Mangroves (wet), Middle Atlantic Bight, Narragansett Bay, Neuse Estuary (early summer 1997), Neuse Estuary (late summer 1997), Neuse Estuary (early summer 1998), Neuse Estuary (late summer 1998), Northern Benguela Upwelling, and Southern New England Bight. The second type consisted of three compartments, containing taxa supported by macrophytes, phytoplankton, or sediment bacteria. Eleven ecosystems showed this type of compartmentalization, including Florida Bay (dry), Florida Bay (wet), Graminoids (dry), Graminoids (wet), St. Marks Seagrass, site 1 (January), St. Marks Seagrass, site 1 (February), St. Marks Seagrass, site 2 (January), St. Marks Seagrass, site 2 (February), St. Marks Seagrass, site 3 (January), St. Marks Seagrass, site 4 (February), and Sylt-Romo Bight. The third type of compartmentalization separated aquatic from terrestrial communities, corresponding to just two ecosystems: Cypress (dry) and Cypress (wet).

We detected between two and five compartments in each of the 28 food webs. Over 80% of them yielded modularity that was significantly greater than would be expected by chance alone (Table 1). The modularity Q varied from 0.02 to 0.39, while the modularity values in Monte Carlo simulations varied from −0.006 to 0.010.

In the node deletion experiment (Figure 3), the within-compartment RE exhibited higher value (0.030 ± 0.004; mean ± SEM) than between-compartment RE (0.015 ± 0.003), and the difference is significant (paired t test: t_{27} = 3.727, p < .001).

Our algorithm based on energy channels generated the highest modularity (0.150 ± 0.017; mean ± SEM) across the 28 food webs, compared with algorithms of edge betweenness (0.119 ± 0.020) and random walk (0.137 ± 0.019). Different algorithm generated significantly different modularity (LME: F_{2,54} = 3.584, p = .035; Figure 4). Here, our algorithm generated significantly higher modularity than the algorithm of edge betweenness (Tukey test: z = 2.667, p = .023), but not the algorithm of random walk (Tukey test: z = 1.128, p = .778).

4 | DISCUSSION

Most food webs in our study were found to be strongly compartmentalized, in agreement with previous research (Guimera et al., 2010; Krause et al., 2003; Rezende et al., 2009). The compartmentalization of food webs has been claimed to arise from subhabitats within the environment (Krause et al., 2003) or because of phylogenetic patterns within the community (Cattin, Bersier, Banašek-Richter, Baltensperger, & Gabriel, 2004; Rezende et al., 2009). Here, we demonstrated that energy channels could be another origin of compartmentalization. Being compartmentalized is to a community’s advantage as compartments act to buffer the effects of perturbation.

Energy channels derived from producers and detritus (namely “green” and “brown” food webs) have been investigated in recent studies (Butler et al., 2008; Rooney & McCann, 2012; Rooney et al.,
Most of these studies focused on the coupling of the two channels, which was believed to increase the stability of ecosystems. In these studies, phytoplankton and detritus were chosen as the origin of energy, and each food web was divided into two subwebs based on their energy origin; however, this might be an oversimplification. Based on our results, different types of detritus may sustain very different compartments. For example, suspended detritus and sediment detritus sustained different compartments in most of our 28 ecosystems. On the other hand, phytoplankton may sustain the same compartment with some types of detritus (e.g., suspended detritus or dissolved organic carbon). Using a more elaborate partition of energy channels may provide new insights into studying the coupling of channels.

Across all 28 food webs in our study, the number of compartments varied from two to five. This is highly consistent with prior studies, where these numbers varied from one to six across 17 food webs (three of them overlap with ours; Krause et al., 2003) or from two to three across nine food webs (one of them overlap with ours; Gauzens, Thébault, Lacroix, & Legendre, 2015). However, such small number did not provide support for the original argument of complex system
community detection algorithms in complex network studies, that is, algorithms based on edge betweenness (Newman & Girvan, 2004) and random walk (Pons & Latapy, 2005). This shows the high feasibility of our algorithm and the high compartmentalization of different channels, especially considering that the other two algorithms use optimization to maximize the value of modularity.

Based on our results, several pairs of energy channels could be regarded as the compartments of a food web, including pelagic versus benthic channels, macrophyte-derived versus phytoplankton-derived channels, bacteria-derived versus algae-derived channels, and terrestrial (i.e., leaf/wood-derived) versus aquatic channels. This is not in conflict with the argument that compartments arise from subhabitats (Krause et al., 2003; Pimm, 1991). Each habitat requires a set of adaptations from its component species, which may preclude a large number of interactions between species in different habitats and encourage more within-habitat interactions (Pimm, 1991), that is, creating habitat-based compartments. Meanwhile, different habitats can sustain different energy channels, for example, pelagic versus benthic and terrestrial versus aquatic. Moreover, there can still be different compartments within the same subhabitat, for example, benthic bacteria and benthic phytoplankton, and sometimes different subhabitats may share the same energy channel, depending on the scale of subhabitats. This suggests that energy channels and subhabitats may both contribute to the compartmentalization of food webs. Besides, our results appear to emphasize the physical separation between benthic and pelagic habitats. This implies the limitation of our datasets: All the 28 food webs we studied are aquatic or at least contain aquatic part. This to some extent limits the application of our conclusions, although it would be interesting to apply our algorithm to terrestrial or other food webs and compare the results with our conclusions in this study.

Basal resources (producers and detritus) often exhibit quite variable traits, such as habitat type and body size, which can attract very different types of primary consumers (Rooney & McCann, 2012). This characteristic supports diverse energy channels, which are in turn coupled by mobile predators at higher trophic levels. The partition of food webs in our study suggests that this coupling is not so strong and that the channels can be highly independent of one another. Therefore,
potential perturbations should be constrained within a single compartment, as suggested in previous studies (Krause et al., 2003; Stouffer & Bascompte, 2011). This may be a vital stabilizing characteristic of natural ecosystems in the face of accelerating biodiversity loss (Barnosky et al., 2011).

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CONFLICT OF INTERESTS

We have no competing interests.

DATA ACCESSIBILITY

All data used in this paper are available in the “enaR” package (Borrett & Lau, 2014) in R (https://cran.r-project.org/web/packages/enaR/index.html).

AUTHORS’ CONTRIBUTIONS

HZ and LZ were responsible for research design. LZ, XL and WT performed numerical simulation and statistical analyses. LZ drafted the main text and prepared the figures. All authors were involved in discussions and editing.

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