Time-Stratified Case Crossover Study of the Association of Outdoor Ambient Air Pollution With the Risk of Acute Myocardial Infarction in the Context of Seasonal Exposure to the Southeast Asian Haze Problem

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Background—Prior studies have demonstrated the association of air pollution with cardiovascular deaths. Singapore experiences seasonal transboundary haze. We investigated the association between air pollution and acute myocardial infarction (AMI) incidence in Singapore.

Methods and Results—We performed a time-stratified case-crossover study on all AMI cases in the Singapore Myocardial Infarction Registry (2010–2015). Exposure on days where AMI occurred (case days) were compared with the exposure on days where AMI did not occur (control days). Control days were chosen on the same day of the week earlier and later in the same month and year. We fitted conditional Poisson regression models to daily AMI incidence to include confounders such as ambient temperature, rainfall, wind-speed, and Pollutant Standards Index. We assessed relationships between AMI incidence and Pollutant Standards Index in the entire cohort and subgroups of individual-level characteristics. There were 53,948 cases. Each 30-unit increase in Pollutant Standards Index was association with AMI incidence (incidence risk ratio [IRR] 1.04, 95% CI 1.03–1.06). In the subgroup of ST-segment–elevation myocardial infarction the IRR was 1.00, 95% CI 0.98 to 1.03, while for non–ST-segment–elevation myocardial infarction, the IRR was 1.08, 95% CI 1.05 to 1.10. Subgroup analyses showed generally significant. Moderate/unhealthy Pollutant Standards Index showed association with AMI occurrence with IRR 1.08, 95% CI 1.05 to 1.11 and IRR 1.09, 95% CI 1.01 to 1.18, respectively. Excess risk remained elevated through the day of exposure and for >2 years after.

Conclusions—We found an effect of short-term air pollution on AMI incidence, especially non–ST-segment–elevation myocardial infarction and inpatient AMI. These findings have public health implications for primary prevention and emergency health services during haze. (J Am Heart Assoc. 2019;8:e011272. DOI: 10.1161/JAHA.118.011272.)

Key Words: myocardial infarction • population • haze • Singapore • air pollution

The role of ambient air pollution in the pathogenesis of a diverse range of acute and chronic diseases is increasingly recognized.1 Southeast Asian (SEA) transboundary haze because of forest fires is a major public health problem, exacting a large economic and health toll on the region. The Global Burden of Disease Study identified fine particulate matter (PM) in outdoor air to be the ninth leading risk factor for disease worldwide,2 while the World Health Organization attributes 1 in every 8 deaths to air pollution.3 There is growing epidemiological evidence that air pollution contributes a heterogeneous and currently poorly understood, yet important role in a range of health outcomes ranging from low birth weight to sudden cardiac death. The ubiquitous and involuntary exposure to air pollution makes it a formidable and highly relevant preventive medicine challenge.

Cardiovascular disease is the leading cause of death worldwide.4 A 2010 update to a scientific statement from the American Heart Association concluded that the overall absolute risk for mortality because of particulate matter exposure is greater for cardiovascular disease compared with pulmonary disease after both short- and long-term exposures, and that the evidence for association of PM exposure was
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Clinical Perspective

What Is New?

- This nationwide registry study using a time-stratified case-crossover design found a transient effect of short-term air pollution on acute myocardial infarction incidence, especially non-ST-segment–elevation myocardial infarction and inpatient acute myocardial infarction, after adjusting for other meteorological indicators and stratifying by individual characteristics.

What Are the Clinical Implications?

- These findings have public health implications for acute myocardial infarction prevention and emergency health services delivery during haze, in particular, the Southeast Asian haze problem.

moderate for ischemic heart disease, heart failure, and ischemic stroke, and mixed for arrhythmia and cardiac arrest. Acute myocardial infarction (AMI) is the main acute presentation of ischemic heart disease where there is irreversible myocardial ischemia leading to significant morbidity and mortality; therefore, understanding the relationship between ambient air pollution and AMI is of tremendous public health interest.

Short-term increase in air pollution has previously been shown to be capable of triggering acute coronary syndromes. In a large comparative risk assessment study of the triggers of myocardial infarction, amongst all the potential triggers studied, the highest population attributable fractions (which comprise both the effect size and prevalence of a risk factor), were contributed by road traffic and concentration of PM with aerodynamic diameter <10 μm (PM_{10}). These studies varied in the accuracy, completeness, and representativeness of both the pollution and disease data. Further, unanswered questions remain about the presence of susceptible populations, possible lag effects, and relevance in the context of the Southeast Asian (SEA) transboundary haze problem.

The SEA haze problem refers to periodic episodes of transboundary and large-scale air pollution episodes that have been recorded since 1972. These events have exacted adverse health and economic impact in Indonesia, Singapore, Malaysia, Brunei Darussalam, southern Thailand, northern Laos, and as far as the southern Philippines. Haze crises such as those that occurred in 1997, 2006, 2013, and 2015 have damaging effects on tourism, transport, food and water quality, urban and rural livelihood, and overall human productivity. Unlawful industrial-scale slash-and-burn land-clearing agricultural practices in the region have been incriminated for sparking off forest fires which releases acrid smoke, dust, and particulate matter into the atmosphere. The haze situation surfaces recurrently in Singapore, usually coinciding with the dry season from July to September, when the southwest monsoon also shifts haze towards Singapore. These episodes have been implicated in damages amounting to an estimated of US $4.5 billion for the fire episodes in 1997 alone. These included health impacts, reduced crop yield, preventive expenditures, accidents, loss of life, evacuations, and the loss of confidence of foreign investors. In 2002, member states of the Association of Southeast Asian Nations ratified the Association of Southeast Asian Nations Agreement on transboundary haze pollution to monitor, prevent, and mitigate transboundary haze through international cooperation.

Singapore is a small, densely-populated island city-state situated in SEA, and experiences periodic large-scale transboundary haze originating from the region. It is hence susceptible to wide day-to-day fluctuations in ambient air pollutant levels over decades. Singapore has robust surveillance capabilities for air pollution and other environmental parameters. These characteristics make Singapore an ideal natural population laboratory to study short-intermediate term health impacts arising from the SEA haze problem.

The objective of this study is to investigate the association between ambient air pollution and AMI occurrence using a time-stratified case-crossover design while adjusting for meteorological parameters and stratifying by individual characteristics. It is hypothesized that exposure to increased Pollutant Standards Index (PSI) is associated with an increase in the number of AMI cases. Other research questions are whether the effect is highest on same day of exposure or after lagged terms of a few days, and whether the risk differs between various subgroups. Findings would inform public health policies relating to measures to reduce air pollutants as well as those to mitigate their effect on susceptible subgroups of the population.

Methods

Data, methods used in the analysis, and materials used to conduct the research available will not be made available to other researchers for purposes of reproducing the results or replicating the procedure. Exposure data are available to the public from the respective agency’s website, and disease data are owned by National Registry of Diseases Office and data sharing is limited by their policies.

Setting

Singapore is a heavily urbanized island city-state in Southeast Asia with a population of 5.5 million over a land area of 719.1 square kilometers. Singapore has a gross domestic
product of 295.7 billion dollars and a life expectancy of 82.1 years. Singapore lies 1.5° north of the equator and is located at the end of the Malayan Peninsula. Its climate has been classified as tropical rainforest (Köppen-Geiger classification system). As a result of its geographical location and maritime exposure, its climate is characterized by uniform temperature and pressure, high humidity, abundant rainfall, and no true distinct seasons. Tertiary health care was delivered by 7 public general hospitals and several private hospitals, of which 5 provided around-the-clock emergency PCI service.18

From prior studies, the age-standardized incidence rate of STEMI in Singapore was 56.6 per 100 000 population from 2010 to 2012.19 Median first medical contact-to-door time was 33.5 minutes20 while median door-to-balloon time was 64 minutes.19 Around 50% of STEMI cases in Singapore presented to the hospitals via ambulances.19

Study Population and Outcome Data—The Singapore Myocardial Infarction Registry

The SMIR (Singapore Myocardial Infarction Registry) is a nationwide registry managed and funded by the National Registry of Diseases Office, Ministry of Health Singapore,21 and collects epidemiological and clinical data on AMI cases diagnosed in all public and private sector hospitals and a small number of out-of-hospital AMI deaths certified by medical practitioners in Singapore. AMI is a disease whose notification to the registry has been mandated by the National Registry of Diseases Act enacted in 2012. Public sector cases comprise 98% of the registered cases.

Registry data were received from various sources on a monthly basis and were processed monthly to obtain unique cases. The sources of data included patient medical claim listings, hospital in-patient discharge summaries, cardiac biomarker listings from hospital laboratories and the national death registry. The International Classification of Diseases, Ninth Revision, Clinical Modification (ICD-9-CM) code 410 was used to identify AMI cases diagnosed from 2010 to 2011 while ICD-10 (Australian Modification) codes I21 and I22 were used for AMI cases diagnosed in 2012. The differentiation between ST-segment–elevation myocardial infarction (STEMI) and non-STEMI (NSTEMI) was based on presenting symptoms, cardiac biomarkers and ECG assessment, and aligned with clinician’s diagnosis documented in the physical case notes and electronic medical records. STEMI was defined as follows: typical chest pain of 30 minutes and significant ST segment elevation (>0.1 or 0.2 mV on 2 adjacent limb or precordial leads, respectively, or new left bundle-branch block) and confirmed subsequently by a rise in biomarkers. All ECGs were interpreted, and all diagnoses were adjudicated centrally at the National Registry of Diseases Office. The multinational monitoring of trends and determinants in cardiovascular disease (MONICA) criteria22 were used for defining episodes.

Detailed patient data were extracted from clinical medical records including ambulance records, Emergency department notes, clinical charts, and discharge summaries, by dedicated registry coordinators from the SMIR. Yearly audits on data collected were done to ensure data accuracy and inter-rater reliability of at least 95%.

The primary outcome variable of our study is the occurrence of an AMI. We considered all cases of AMI in Singapore from 2010 to 2015. Recurrence of AMI after 28 days of a recorded AMI episode was considered a separate episode.

The event date was taken to be the date of onset as we included both inpatient and outpatient AMI. Although the date of onset may be subjected to recall bias by patient, it was checked against the start of acute symptoms, serial changes in ECG, elevation in cardiac biomarkers, and fatal collapse, whichever was earliest.

Environmental Data

The primary exposure was 24-hour average PSI. PSI is an air quality index used to indicate the level of pollutants in the air. This was based on a scale devised by the United States Environmental Protection Agency to provide a way for news agencies to report daily air quality. PSI has been used in several countries including the United States, Brunei Darussalam, and Singapore. The National Environment Agency in Singapore classifies 24-hour PSI into ranges of good (0–50), moderate (51–100), unhealthy (101–200), very unhealthy (201–300), and hazardous (>300).23

PSI is computed based on 6 air pollutants: fine particulate matter with aerodynamic dynamic smaller than 2.5 μm (PM_{2.5}), PM_{10}, sulfur dioxide (SO_{2}), carbon monoxide (CO), ozone (O_3), and nitrogen dioxide (NO_2). For each pollutant, a sub-index is calculated from a segmented linear function that transforms ambient concentrations onto a scale extending from 0 through 500.24 PSI is then computed to be the maximum of the 6 sub-indices. PM_{2.5} is the major pollutant released by forest fires,25 and the World Health Organization guideline level for 24-hour mean PM_{2.5} is 25 μg/m^3.26

In Singapore, ambient pollutant levels are continuously monitored at >20 telemetric air quality monitoring stations across the island.27 Exposure data were retrieved from local government websites. The data were those that are made publicly available by the government agencies but required the authors to write a script to aggregate the data into a usable format. Historical 24-hour PSI data were obtained from www.haze.gov.sg which is maintained by the National Environment Agency. Meteorological parameters were obtained.
from www.weather.gov.sg which is maintained by the Meteorological Service Singapore, and included total daily rainfall, daily highest rainfall (over 30-, 60- and 120-minute intervals), daily temperature (mean, maximum, minimum) and daily wind speed (mean, maximum). Data from 5 meteorological stations (Jurong West, Khatib, Simei, Kampong Bahru, and Upper Thomson) scattered across Singapore were obtained.

**Ethics Approval**

The Centralized Institutional Review Board and Domain Specific Review Board granted approval for this study with a waiver of patient consent (CIRB reference number: 2017/2380).

**Statistical Analysis**

Data analysis was performed using Stata version 13.28 Categorical and continuous data were presented as frequency with percentage and median with interquartile range (IQR), respectively. Statistical significance was set at \( P < 0.05 \).

This study used a time-stratified design to control for time trend and other short-term varying confounders like ambient temperature since it compares exposure levels between same weekdays within each month of each year.29 Exposure on days where AMI occurred (case days) were compared with the exposure on days where AMI did not occur (control days). Control days were chosen on the same day of the week earlier and later in the same month in the same year. Daily AMI counts approximately followed Poisson distribution.

We fitted a conditional Poisson regression model to daily AMI incidence that included 24-hour average PSI and potential confounders such as daily average temperature, total rainfall, and average wind speed. All models were adjusted for over-dispersion and autocorrelation, except for a sensitivity analyses without adjustment for over-dispersion and autocorrelation. Based on National Environment Agency’s recommended range, PSI was categorized into 3 categories: good, moderate, and unhealthy (cutoffs previously stated). We assessed the relationship between AMI incidence and PSI range in the entire cohort and in subgroups of demographic and clinical characteristics, which were determined a priori. The subgroups are chosen to allow identification of susceptible subpopulations.

We also investigated percent excess risk of AMI associated with each 30-unit increase in PSI values on the day of incidence (lag 0 day) and subsequent days before the incidence, until there was no further persistent lag effect. The results were presented as incidence risk ratio (IRR) and 95% CI for PSI range and percent excess risk in AMI per 30-unit increase in PSI values using the formula (risk ratio – 1) × 100.

**Results**

**Study Population**

There were 53,948 cases of AMI between 2010 and 2015 qualified for analysis. The median age was 68.9 years (IQR 58.0–79.6) and 65.1% were male. There were 64.3% were non-ST-segment–elevation myocardial infarction (NSTEMI) while 25.0% were STEMI, 24.1% occurred while inpatient, and 81.6% survived to hospital discharge. The characteristics of included cases are further described in Table 1.

**Description of Exposure Data**

Summary characteristics of air pollution and meteorological parameters are shown in Table 2. During the study period,
median daily 24-hour average PSI was 32.8 (IQR 25.7–47.0). Median daily total rainfall on the days that rained was 4.7 mm (IQR 1.0–12.8). Median daily average temperature was 27.7°C (IQR 26.9–28.4). Median daily average wind speed was 7.0 km/h (IQR 6.0–8.5).

Association of 30-Unit Increments in Pollutant Standards Index With Occurrence of AMI

Figure shows the distribution of 3-weekly AMI incidence with weekly average measure of 24-hour PSI. Smoothing of data included 1 lagged term, 1 forward term, and the current observation in the time-series moving average filter.

When considering PSI in terms of 30-unit increments, after adjusting for temperature, rainfall, and wind speed, PSI was significantly associated with increased AMI occurrence, with each 30-unit increment in PSI being associated with IRR of 1.04 (95% CI 1.03–1.06) in the entire cohort (Table 3).

In subgroup analyses of demographic and clinical characteristics, the association remained generally significantly positive, except in the subgroup of STEMI (IRR 1.00, 95% CI 0.98–1.03). In addition, even though the association between PSI and AMI were significant in both subgroups of inpatient and outpatient AMI, the strength of association was higher in the inpatient subgroup compared with the outpatient subgroup, with IRR 1.13 (95% CI 1.09–1.16) and IRR 1.02 (95% CI 1.00–1.04), respectively.

Besides STEMI versus NSTEMI and inpatient versus outpatient, there were no significantly increased susceptibility observed in other subgroups of age, sex, ethnicity, history of

| Subgroups | Incidence Rate Ratio (95% CI) | P Value |
|-----------|------------------------------|---------|
| Entire cohort | 1.04 (1.03–1.06) | <0.001 |
| Without overdispersion and autocorrelation | 1.04 (1.03–1.06) | <0.001 |
| Subgroups | | |
| Age | | |
| <65 y | 1.04 (1.02–1.07) | <0.001 |
| ≥65 y | 1.05 (1.03–1.07) | <0.001 |
| Sex | | |
| Male | 1.06 (1.04–1.08) | <0.001 |
| Female | 1.04 (1.01–1.06) | 0.005 |
| Ethnicity | | |
| Chinese | 1.05 (1.03–1.07) | <0.001 |
| Malay | 1.05 (1.02–1.08) | 0.002 |
| Indian | 1.04 (1.01–1.08) | 0.014 |
| Subtype | | |
| STEMI | 1.00 (0.98–1.03) | 0.940 |
| NSTEMI | 1.08 (1.05–1.10) | <0.001 |
| History of MI/CABG/PCI | | |
| Yes | 1.05 (1.03–1.08) | <0.001 |
| No | 1.05 (1.03–1.07) | <0.001 |
| History of diabetes mellitus | | |
| Yes | 1.06 (1.04–1.08) | <0.001 |
| No | 1.05 (1.03–1.07) | <0.001 |
| History of hypertension | | |
| Yes | 1.05 (1.03–1.07) | <0.001 |
| No | 1.06 (1.04–1.09) | <0.001 |
| History of hyperlipidemia | | |
| Yes | 1.05 (1.03–1.08) | <0.001 |
| No | 1.05 (1.02–1.07) | <0.001 |
| Current/former smoker | | |
| Yes | 1.04 (1.02–1.07) | <0.001 |
| No | 1.06 (1.04–1.08) | <0.001 |
| Place of MI onset | | |
| Inpatient | 1.13 (1.09–1.16) | <0.001 |
| Outpatient | 1.02 (1.00–1.04) | 0.029 |

Incidence rate ratios were estimated using conditional poisson regression adjusted for overdispersion and autocorrelation, with average Pollutant Standards Index, rainfall, temperature and wind speed as random effects covariates. CABG indicates coronary artery bypass grafting; MI, myocardial infarction; NSTEMI, non-ST-segment-elevation myocardial infarction; PCI, percutaneous coronary intervention; STEMI, ST-segment-elevation myocardial infarction.

Figure. Distribution of weekly occurrence of myocardial infarction with Pollutant Standards Index. Smoothing of data included 1 lagged term, 1 forward term, and the current observation in the time-series moving average filter. MI indicates myocardial infarction; PSI, Pollutant Standards Index.
cardiac disease, diabetes mellitus, hypertension, hyperlipidemia, and smoking history.

**Association of Good, Moderate and Unhealthy Ranges of Pollutant Standards Index With Occurrence of AMI**

When considering PSI in terms of categories of good, moderate, and unhealthy ranges, after adjusting for temperature, rainfall and wind speed, taking the good PSI range as a reference, moderate and unhealthy PSI ranges were significantly associated with increased AMI occurrence (Table 4), with IRR 1.08 (95% CI 1.05–1.11) and IRR 1.09 (95% CI 1.01–1.18), respectively.

The associations between moderate and unhealthy ranges of PSI with AMI in the STEMI subgroup were not statistically significant, IRR 1.18 (95% CI 0.89–1.52), respectively.

Excess Risk at Different Lag Terms

In addition, after adjusting for temperature, rainfall, and wind speed, each 30-unit increment in PSI values on the same day (lag 0) and previous 1 day up until 2 years and 4 months (not cumulative) was significantly associated with increased risk of AMI incidence (Table 5), presented at intervals of 2 months.

The excess risks from 30-unit increment in PSI for same day exposure was 4.38 (95% CI 2.66–6.12) and for the 427th day after was 4.76 (95% CI 2.34–7.25) (Table 5).

Limitations

There are several limitations to this study. First, the study design does not allow a causative relationship between PSI and AMI to be proven, even though strongly suggested on the basis of plausibility, temporality, and dose-response.

Secondly, there are possible residual confounding factors because the study design did not permit controlling for behavioral changes related to air pollution, as people take different approaches to mitigation based on their beliefs and attitudes. For example, while the elderly can mitigate the haze by staying at home, those who had to work outdoors during the haze period have no such recourse. Also, those with higher socioeconomic status may have access to air-conditioning and respirator masks (which are proven to improve a range of cardiovascular health measures in patients with coronary disease) which would modify their risk. This same effect may explain findings in other studies that the adverse effects from PM2.5 exposure were higher for individuals with lower education level.

Thirdly, PSI, which is a composite air quality index, was used instead of concentrations of individual constituent pollutants. Therefore, we are unable to examine the relative excess risk contributed by each pollutant. On the other hand, PSI is an easily recognizable indicator for the public to understand, monitor, and interpret. The findings of this study are hence easily actionable for health education of the public and to calibrate as public policy. It is also suspected that the individual pollutants may interact in complex ways that may be not accounted for by the statistical methods used. There is therefore value in considering air pollution in phenotypic groups of their origins, such as forest fire or urban exhaust, rather than in their constituent components. This is supported by a study of 6 cities in the United States showing that for the same particle size and concentration, combustion-origin particles are associated with increased mortality, but not crustal-origin particles. We have previously demonstrated correlation of PSI with other health outcomes such as sudden cardiac death and acute ischemic stroke.

Fourthly, the small number of events incurring on days with PSI in the unhealthy range (n=41) is small and hence the power to detect significant association is reduced for the comparison of unhealthy PSI versus good PSI ranges.

There is also potential for misclassification of AMI onset date as it is partially affected by recall bias by the patient. Also, some symptoms preceding admission may have been angina episodes and not attributed to the index AMI episode.

**Discussion**

This study demonstrated that in a population exposed to wide fluctuations in ambient air pollution, exposure to increases in a composite air pollution index was associated with a markedly increased short-intermediate risk of AMI. The excess risk was most pronounced on the day of the exposure, but remained elevated for the next 5 days. This is, to our knowledge, the first study linking the Southeast Asian haze to AMI, other than indirectly, in our previous study of the effect of PSI on out-of-hospital cardiac arrest (currently in press).

The finding of a significant association of exposure to air pollution and occurrence of AMI is generally consistent with previous studies. We have found 8% and 9% excess risk of AMI when PSI is in the moderate and unhealthy range,
respectively. Furthermore, when considering only NSTEMIs, which seems to be driving the association, the magnitude of the excess risk is greater at 15% and 11%, respectively. Given that practically the entire population is exposed involuntarily to this risk factor, the collective health burden is magnified.

### Table 4. Estimated Incidence Rate Ratio of MI for Each PSI Group for the Entire Study Cohort and by Subgroups of Demographic and Clinical Characteristics (n=2191 Days)

|                          | Good PSI (n=1721) | Moderate PSI (n=429) | Unhealthy PSI (n=41) |
|--------------------------|-------------------|----------------------|----------------------|
|                          | IRR (95% CI)      | P Value              | IRR (95% CI)         | P Value              |
| Entire cohort            | 1.00 (reference)  |                      | 1.08 (1.05–1.11)     | <0.001               |
|                          |                   |                      | 1.09 (1.01–1.18)     | 0.021                |
| Without overdispersion and autocorrelation | 1.00 (reference) |                      | 1.08 (1.05–1.11)     | <0.001               |
|                          |                   |                      | 1.09 (1.01–1.18)     | 0.018                |
| **Subgroups**            |                   |                      |                      |                      |
| **Age**                  |                   |                      |                      |                      |
| <65 years                | 1.00 (reference)  | 1.08 (1.04–1.12)     | <0.001               | 1.09 (0.99–1.21)     | 0.092                |
| ≥65 years                | 1.00 (reference)  | 1.11 (1.07–1.15)     | <0.001               | 1.10 (1.00–1.21)     | 0.052                |
| **Sex**                  |                   |                      |                      |                      |
| Male                     | 1.00 (reference)  | 1.09 (1.06–1.13)     | <0.001               | 1.11 (1.02–1.21)     | 0.018                |
| Female                   | 1.00 (reference)  | 1.11 (1.06–1.16)     | <0.001               | 1.08 (0.96–1.21)     | 0.228                |
| **Ethnicity**            |                   |                      |                      |                      |
| Chinese                  | 1.00 (reference)  | 1.10 (1.06–1.13)     | <0.001               | 1.08 (0.99–1.19)     | 0.070                |
| Malay                    | 1.00 (reference)  | 1.09 (1.03–1.15)     | 0.001                | 1.08 (0.93–1.25)     | 0.306                |
| Indian                   | 1.00 (reference)  | 1.08 (1.02–1.14)     | 0.013                | 1.08 (0.92–1.27)     | 0.331                |
| **Subtype**              |                   |                      |                      |                      |
| STEMI                    | 1.00 (reference)  | 1.02 (0.98–1.07)     | 0.363                | 1.00 (0.89–1.14)     | 0.939                |
| NSTEMI                   | 1.00 (reference)  | 1.15 (1.11–1.18)     | <0.001               | 1.11 (1.01–1.21)     | 0.029                |
| **History of MI/CABG/PCI**|                    |                      |                      |                      |
| Yes                      | 1.00 (reference)  | 1.10 (1.05–1.14)     | <0.001               | 1.04 (0.92–1.18)     | 0.488                |
| No                       | 1.00 (reference)  | 1.10 (1.07–1.14)     | <0.001               | 1.13 (1.04–1.23)     | 0.005                |
| **History of diabetes mellitus** |              |                      |                      |                      |
| Yes                      | 1.00 (reference)  | 1.08 (1.04–1.12)     | <0.001               | 1.10 (0.99–1.22)     | 0.064                |
| No                       | 1.00 (reference)  | 1.12 (1.08–1.16)     | <0.001               | 1.11 (1.01–1.22)     | 0.030                |
| **History of hypertension** |                    |                      |                      |                      |
| Yes                      | 1.00 (reference)  | 1.08 (1.05–1.11)     | <0.001               | 1.09 (1.00–1.19)     | 0.046                |
| No                       | 1.00 (reference)  | 1.14 (1.09–1.19)     | <0.001               | 1.12 (1.00–1.27)     | 0.057                |
| **History of hyperlipidemia** |                 |                      |                      |                      |
| Yes                      | 1.00 (reference)  | 1.09 (1.05–1.13)     | <0.001               | 1.08 (0.98–1.19)     | 0.117                |
| No                       | 1.00 (reference)  | 1.10 (1.06–1.14)     | <0.001               | 1.14 (1.03–1.26)     | 0.015                |
| **Current/former smoker** |                    |                      |                      |                      |
| Yes                      | 1.00 (reference)  | 1.07 (1.03–1.11)     | <0.001               | 1.06 (0.96–1.18)     | 0.250                |
| No                       | 1.00 (reference)  | 1.13 (1.09–1.17)     | <0.001               | 1.12 (1.02–1.24)     | 0.019                |
| **Place of MI onset**    |                   |                      |                      |                      |
| Inpatient                | 1.00 (reference)  | 1.23 (1.17–1.30)     | <0.001               | 1.32 (1.15–1.52)     | ~0.001               |
| Outpatient               | 1.00 (reference)  | 1.05 (1.02–1.08)     | 0.002                | 1.03 (0.95–1.12)     | 0.522                |

Incidence rate ratios were estimated using conditional poisson regression adjusted for overdispersion and autocorrelation, with average PSI, rainfall, temperature and wind speed as random effects covariates. CABG indicates coronary artery bypass grafting; IRR, incidence rate ratio; MI, myocardial infarction; PCI, percutaneous coronary intervention; PSI, Pollutant Standards Index.
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Table 5. Estimated Percent Excess Risk of Myocardial Infarction for Each 30-Unit Increment of Pollutant Standards Index at Different Lag Term (n=2191, d)

| Lag, d | Excess Risk in % (95% CI) | P Value |
|--------|---------------------------|---------|
| 0      | 4.38 (2.66-6.12)          | <0.001  |
| 61     | 5.69 (3.98-7.43)          | <0.001  |
| 122    | 6.17 (3.86-8.53)          | <0.001  |
| 183    | 4.35 (2.02-6.73)          | <0.001  |
| 244    | 5.02 (2.70-7.40)          | <0.001  |
| 305    | 3.53 (1.16-5.95)          | 0.003   |
| 366    | 3.70 (1.32-6.14)          | 0.002   |
| 427    | 4.76 (2.34-7.25)          | <0.001  |
| 488    | 2.12 (–0.54 to 4.84)      | 0.119   |

Lag 0 refers to same day exposure; lag 61 refer to exposure at 61 days prior. Excess risks were estimated using conditional poisson regression adjusted for overdispersion and autocorrelation, with average Pollutant Standards Index, rainfall, temperature and wind speed as random effects covariates.

Table 5. Estimated Percent Excess Risk of Myocardial Infarction for Each 30-Unit Increment of Pollutant Standards Index at Different Lag Term (n=2191, d)

intermediate term risk, and would not capture long-term risk. Therefore, the effect of air pollution on AMI is significantly greater than suggested by this risk estimate magnitude alone.

The finding that only NSTEMI, but not STEMI is associated with increases in PSI is of interest. A case-crossover study of the England and Wales Myocardial Ischaemia National Audit Project, which was the only other study which examined the differential effect on NSTEMI versus STEMI, showed comparable results. In that study of 202 550 STEMI and 322 198 NSTEMI events, air pollutants were significantly associated with NSTEMI but not STEMI. In addition, taking AMI as a whole cohort (including both NSTEMI and STEMI), the excess risk ceased to remain significant. However, other smaller studies on AMI (not designed specifically to look at the differential effect on NSTEMI and STEMI), do not corroborate these findings, with several finding significantly increased risk for STEMI and 1 study showing no increased risk for NSTEMI. Given these mixed findings, it is difficult to postulate the underlying reason for the differential effects on STEMI and NSTEMI, even though knowing that mechanically, STEMI is more often because of an acute plaque rupture than NSTEMI.

Further clues are found in the related finding that excess risk for inpatient AMI is greater than for outpatient AMI. In general, inpatient AMIs more often occur in patients admitted for unrelated diagnoses such as sepsis or surgery and develop AMI as a result of relative myocardial oxygen demand-supply mismatch. This may suggest that the effect of air pollution on AMI is less likely to be mediated through primary plaque rupture processes.

In terms of susceptible sub-populations, we have found no clear evidence of increased susceptibility in elderly patients or patients with cardiac history, diabetes mellitus, hypertension, or hyperlipidemia. This is in general agreement with previous studies. Knowledge of susceptible sub-populations, if any, is important for policy makers in formulating health advisories to target these at-risk populations, and to plan cost-effective interventions.

There is literature supporting a range of possible mechanisms for developing AMI in response to air pollution, and they generally fall into categories of coagulation, inflammation, vascular dysfunction, and autonomic dysfunction. These include experimental cellular, histological, animal, and healthy volunteer studies. It is likely that multiple mechanisms jointly contribute to the phenomenon.

One strength of this study lies in the high-quality outcome data from a national registry featuring comprehensive case capture. Legislated mandatory reporting of AMI and central adjudication of case qualification contributed to comprehensive and consistent case capture. In addition, we used exposure data that are measured directly by stations located around Singapore, and hence there was less exposure misclassification caused by extrapolating to rural areas via modeling. Additionally, the conditional Poisson regression model used in this study accounted for over-dispersion and autocorrelation in the time-dependent counts data.

Millions of people worldwide are exposed to seasonal high levels of air pollution from forest fires, a modifiable risk factor, making it a tremendous public health issue. Being previously shown to be the risk factor that contributes the highest population attributable fraction to myocardial infarction alone, air pollution presents a high yield disease prevention target for public health administrators. This study adds a Southeast Asian context to the growing body of evidence on the effect of air pollution on health, which as a whole, presents a compelling argument for concerted national efforts and intensified international cooperation to develop sustainable programs to tackle the haze problem in Southeast Asia and worldwide. Also, quantification of the health impact helps guide policy makers in evidence-based policy design and resource allocation. These measures may include issuing health advisories, public N95 respirator distribution programs, school closure policies, city planning from the urban development perspective, emergency medical resources deployment, as well as training doctrines for military training institutes.

Conclusion

During the study period, exposure to higher PSI was associated with an increased short-intermediate term risk of AMI in Singapore. Excess risk remained elevated through the day of exposure and for up to 5 days after exposure. This is, to our knowledge, the first study linking the Southeast Asian
haze problem to AMI. These results have public health implications for the region.

Disclosures
None.

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