Research on Threshold Effect of Housing Price on Real Estate Inventory

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Abstract—Through the systematic GMM estimation of the panel data of 35 large and medium-sized cities in China from 2005 to 2016, this paper verifies the nonlinear relationship between housing prices and real estate inventory. On this basis, the threshold regression model is established by modifying the econometric model. It is suggested that cities with high real estate inventory should guide the housing price to return to a reasonable range, while cities with low inventory should increase land supply to strengthen housing security.

Keywords—housing price; real estate inventory; system GMM estimation; threshold regression

I. INTRODUCTION

In the first half of 2019, the news of housing sale in Hegang was widely reported by the media, with the houses being sold at the price of "cabbage", indicating that the city's real estate inventory is under great pressure. As a matter of fact, by December 2018, inventory in 44 of the 100 key cities in China had increased year on year, while in 10 cities, including JingDeZhen, HuiZhou and LangFang, inventory had increased by more than 50% year on year. History always amazing similar, from the point of relevant statistics, China's commercial housing sale area reached 719 million square meters at the end of 2015, and the area of commercial housing for sale in China is still as high as 695 million square meters at the end of 2016, the decline was only 3.2%, thus it can be seen, even if the implementation of the preferential policy that buy a house, high inventory phenomenon still exists, commercial housing inventory to effect is not obvious. Although the housing price and inventory in some cities have returned rationally to a certain extent nationwide, due to the cyclical development of the real estate market itself, inventory in some cities has been gradually high since 2018. Therefore, after the exit of the real estate destocking policy at the national level, it is of certain theoretical and guiding significance to study the relationship between housing price and real estate inventory to observe the cyclical fluctuation of real estate inventory and implement precise real estate regulation.

II. LITERATURE REVIEW

At present, researches on China's real estate inventory are mostly concentrated in China. Due to the limitation of geographical factors, there are few researches on China's real estate inventory abroad. However, foreign studies on real estate inventory mainly focus on the calculation of vacancy rate and how to maintain the vacancy rate within a reasonable range. The United States is the first country in the world to publish the vacancy rate calculation index, while the European Union, Japan and other countries also have different calculation on the vacancy rate (Wen Jianwu et al, 2010). At present, there is no uniform calculation standard and caliber to measure the vacancy rate. Struky et al. (1983) believed that the reasonable vacancy rate range between developed countries and developing countries is different. When the housing vacancy rate is too high, housing price should be adopted to adjust the excess housing.

Domestic theoretical research on real estate inventory mainly focuses on motivation and countermeasure analysis. In terms of the motivation of real estate inventory, Li Xiaoman (2017) mainly discussed the reasons for the coexistence of high inventory and high housing price in some hot second-tier cities. Fang Yao (2016) sorted out the influencing factors of real estate inventory through empirical analysis, including housing price, income and economic growth. Real estate countermeasures to inventory, Wang Xiaoguang (2017) think the destocking of 2016 is unsatisfactory not only failed to reduce inventory but also leverage the real estate market bubble increased, the second-tier cities should focus on inventory to curb speculative demand, while three or four line cities reduce inventory should be combined with a new type of urbanization. Hu Zuquan (2016) believes that speeding up the citizenization of migrant workers is the key to fighting a “destocking” war of annihilation, and in addition, housing price should be reduced and land supply should be improved. Zhou Youman (2016) believes that real estate destocking can be accelerated from the two aspects of real estate enterprise transformation and financial institutions boosting.

The empirical research on real estate inventory in China mainly focuses on the impact of real estate inventory on national economy and the impact of external factors on real estate inventory. In terms of the impact of real estate inventory on the national economy, Liu Xing and Zhang Jing (2017) investigated the impact of the money supply and commodity housing inventory on the housing price, and studied the relationship between the money supply and commodity housing inventory in China and the housing price.
in 12 major cities from 2008 to 2012. Shen Bo (2016) used the spatial panel model to empirically study the impact of real estate destocking on regional financial stability, and the results showed that in real estate inventory, backlogged inventory and short-term inventory had different degrees of harm to financial stability in the eastern and central regions. External factors impact on real estate stocks, such as Liu (2017) used the data of 35 large and medium cities from 2001-2013 empirically the urbanization and lower the prices of real estate to the influence of the inventory, the results show that urbanization and lower prices really can effectively absorb the current our country real estate excess inventories, but the east and west showed obvious regional differences. Zhang Xiekui and Wu Siyi (2016) selected quarterly data of various variables from March 2004 to June 2016 and established a VAR model to investigate the interaction between short-term loan interest rate (long-term loan interest rate), domestic loan scale of developers, broad money supply and national inventory of commercial housing.

Empirical results of existing literatures show that housing price has different effects on real estate inventory. Liu (2017) think housing prices have a positive impact on real estate stocks, Fang Yao (2016) through analyzing the influencing factors of real estate stocks, the results showed that home prices impact on the stock is positive, and Chen Hong and Li Chao (2017) think that the urbanization rate, the commercial housing price, income level and the medium and long term loan interest rate will cause real estate inventory changes, including the increase in the price of commercial housing would reduce real estate stocks. Therefore, it is necessary to further empirically test the impact of housing prices on real estate inventory. In addition, the above literatures start from the linear relationship to study the impact of housing price on real estate inventory, without considering the possible non-linear relationship. The influence degree of housing price on real estate inventory may change with the passing of time. Therefore, it is of great value and significance to study the non-linear influence of housing price on real estate inventory. On the basis of verifying the non-linear relationship between housing price and real estate inventory, this paper takes the per capita disposable income of urban residents as the threshold variable to specifically investigate the change of housing price's influence on real estate inventory with the increase of per capita disposable income of urban residents.

III. THEORETICAL ANALYSIS

A. Analysis of the Impact of Housing Price on Real Estate Inventory

Supply and demand are the most basic market influencing factors, which determine the digestibility of commercial housing inventory. The supply and demand of the real estate market are closely related to the housing price. The housing price will affect both the supply and demand of the real estate market. When the housing price rises, the housing supply will increase while the housing demand will decrease. According to economic theory, the inventory of a good is equal to the supply of the good minus the demand. In this paper, real estate inventory is used to represent the supply and demand of housing, and the following formula is established:

\[ F = S - D \]  

(1)

Since housing price will affect both supply and demand of housing, so both \( S \) and \( D \) are functions of \( p \), the above equation can be rewritten as:

\[ F(p) = S(p) - D(p) \]  

(2)

In order to study the impact of housing price on real estate inventory, the partial derivative of \( p \) is obtained as follows:

\[ \frac{dF(p)}{dp} = \frac{dS(p)}{dp} - \frac{dD(p)}{dp} \]  

(3)

When the housing price rises, \( \frac{dS(p)}{dp} \) is positive, while \( \frac{dD(p)}{dp} \) is negative, \( \frac{dF(p)}{dp} \) must be positive, so real estate inventory should increase with the rise of housing price, and they are positively correlated. The rise of housing price stimulates the supply of housing by real estate developers, while the high housing price restrains the demand of residents of houses. Therefore, theoretically, housing prices should have a positive impact on real estate inventory.

B. Analysis of the Non-linear Relationship Between Residents' Disposable Income and Inventory Through Housing Price

Domestic and foreign scholars have shown that housing price and residents' disposable income are the most important factors influencing housing demand. Theoretically, the increase of residents' disposable income will stimulate residents' demand for house purchase. That is to say, under the assumption that housing price and other commodity prices remain unchanged, housing consumption will increase with the increase of income, \( \Delta D_1 \) is positive. When the housing price changes, there will be both substitution effect and income effect. The rise of housing price will inhibit the purchase demand of residents and produce a negative substitution effect (denoted by \( \Delta D^3 \)). Moreover, since housing is a normal commodity, the rise of housing price will cause the real income of residents to decrease, thus the consumption of housing will decrease, that is, negative income effect will be generated simultaneously (denoted by \( \Delta D^4 \)):

\[ \Delta D_2 = \Delta D^3 + \Delta D^4 \]  

(4)

When residents' disposable income and housing prices rise at the same time, whether the total demand for housing will increase depends on the size of \( \Delta D_1 \) and \( \Delta D_2 \). The speed of inventory increase is affected by both income and housing price. Although both income and housing price have
risen in recent years, the increase of income is far less than the speed of housing price rise.

According to the above theoretical analysis on the relationship between housing price and real estate inventory, although the rise in housing price will lead to the increase in inventory, the positive correlation between of them may be affected by the increase in residents’ disposable income. With the increase of residents’ disposable income, the degree of positive impact of housing price on real estate inventory will change, and there may be a non-linear relationship between housing price and inventory.

Based on the above theoretical analysis, this paper proposes the following hypothesis:

\[ \text{Real estate inventory} = \beta_0 + \beta_1 \text{Real estate inventory}_{t-1} + \beta_2 \text{Housing price}_{t-1} + \beta_3 \text{Per capita disposable income}_{t-1} + \gamma_t + \mu_t + \xi_t \]

(5)

Research hypothesis 1: The increase of real estate inventory depends on the rise of housing price, and housing price will have a non-linear effect on real estate inventory.

Research hypothesis 2: With the increase of residents’ disposable income, the positive impact of housing price on real estate inventory is increasing.

IV. EMPirical AnalYsis

A. Data Sources and Unit Root Test

In this paper, the available area of commercial housing, housing price, per capita disposable income of urban residents, housing completed area, population density, 3-year bank loan interest rate and other statistical data of 35 large and medium-sized cities in China from 2005 to 2016 are used. Among them, the interest rate is a national indicator, reflecting the change of capital cost. The data structure of each variable is shown in “Table I”:

| Variables | Meaning | Sample Size | Mean Value | Standard Deviation | Minimum | Maximum |
|-----------|---------|-------------|------------|--------------------|---------|---------|
| fsa       | Real estate inventory | 420        | 3142251    | 3171545            | 133600  | 206459  |
| p         | House prices          | 420        | 5904.02    | 3840.97            | 1552    | 24723   |
| pcdi      | Per capita disposable income | 420 | 23846.31  | 17874.05           | 7990    | 319421  |
| ls        | Land supply           | 420        | 8790792    | 7816949            | 609496  | 39906286 |
| lr        | The interest rate     | 420        | 6.02       | 0.68               | 4.75    | 7.56    |
| dp        | The population density| 420        | 644.19     | 402.71             | 123.95  | 2269.23 |

In order to avoid false regression, it is necessary to perform unit root test on the data involved in the housing demand measurement model before regression analysis of the model. Since the data used in this paper are panel data, it is necessary to perform panel unit root test on the data. Generally, the panel unit root is tested in the following four ways: LLC test, IPS test, Fisher-Dfuller (ADF) test and Fisher Phillips-Perron (PP) test. LLC test was homogeneous panel unit root test, while IPS test, Fisher-Dfuller (ADF) test and Fisher Phillips-Perron (PP) test were heterogeneous panel unit root test. In order to avoid heteroscedasticity problems in subsequent regression analysis, all variables were converted to natural logarithm form before the unit root test. The test results show that the sequence of horizontal values of all variables is stable and meets the requirements of modeling. The panel unit root test results of \( \ln p \) and \( \ln fsa \) are shown in “Table II”:

| Variables | Levin-Lin | IPS | Fisher-ADF | Fisher-PP |
|-----------|-----------|-----|------------|-----------|
| \( \ln fsa \) | -18.6494*** (0.0000) | -7.1205*** (0.0000) | 7.4077*** (0.0000) | 5.6573*** (0.0000) |
| \( \ln p \)   | -14.0097*** (0.0000) | -2.9221*** (0.0017) | 23.1891*** (0.0000) | 4.3090*** (0.0000) |

Note: P values are shown in brackets. “***”, “**” and “*” mean significant at 1%, 5% and 10% levels, respectively.

1) Testing the non-linear relation between housing price and inventory

Theoretically, indexes that can affect the sales area of commercial housing and the area for sale will have a certain

\[ \ln fsa_i = \beta_0 + \beta_1 \ln fsa_{i-1} + \beta_2 \ln p_i + \beta_3 C_i + \gamma_t + u_i + \xi_i \]

Where, \( \ln fsa_i \) and \( \ln fsa_{i-1} \) respectively represent the area of sales for the city’s commercial housing in the period of \( i \) and the period of \( i-1 \); \( \ln p_i \) is the housing price of city \( i \) in the period \( i \); \( C_i \) is the combination of control variables; \( \gamma_t \) is the time fixed effect; \( u_i \) is the individual fixed effect; \( \xi_i \) is the random error term.

In order to verify whether there is a linear relationship between housing price and inventory generation, this paper...
adds the square term on the basis of equation (5) and establishes the following model:

\[
\lnfsa_{it} = \beta_0 + \beta_1 \lnfsa_{it-1} + \beta_2 \lnp_{it} + \beta_3 \lnp^2_{it} + \beta_4 \lnpcdi_{it} + \gamma_i + u_i + \xi_{it} \tag{6}
\]

At the same time, in order to investigate the adjustment effect of household income on housing price and inventory, we introduced the interaction item between housing price level and per capita disposable income of urban residents on the basis of equation (5). Thus, the following model was established:

\[
\lnfsa_{it} = \beta_0 + \beta_1 \lnfsa_{it-1} + \beta_2 \lnp_{it} + \beta_3 \lnp^2_{it} + \beta_4 \lnpcdi_{it} * \lnpcdi_{it} + \beta_5 C_{it} + \gamma_i + u_i + \xi_{it} \tag{7}
\]

In this paper, equations (1) - (3) are respectively estimated by the generalized moment estimation method. System of generalized moment estimation is improved on the base of generalized moment estimation method can effectively solve the dynamic change of explanatory variables and the endogenous problem of explanatory variables in the model (Richard and Stephen, 1998). It can also control area fixed effect and annual fixed effect. In view of the revised two-step system GMM is more robust than step system, this paper uses a two-step system GMM estimation. In order to eliminate the impact of inflation on commodity housing price, urban residents’ per capita disposable income and interest rate, we use consumer price index to convert commodity housing price. Urban residents’ per capita disposable income and interest rate into fixed base price index (based on 2005) to get the actual amount. In addition, considering the possible multicollinearity of the square term and the interaction term in equations (5) and (6), in order to ensure the significance of the original variable coefficients, the constructed square term and the interaction term were decentralized in this study to ensure the reliable estimation results.

The estimated results of the above three formulas are shown in “Table III”:

| Variables          | Model(1) | Model(2) | Model(3) |
|--------------------|----------|----------|----------|
| \(\lnfsa_{it-1}\)  | 0.734**  | (0.016)  | 0.677**  | (0.022)  | 0.843**  | (0.020)  |
| \(\lnp_{it}\)      | 0.231**  | (0.020)  | 0.260**  | (0.025)  | 0.277**  | (0.031)  |
| \(\lnp^2_{it}\)    | 0.127**  | (0.017)  |          |          |          |          |
| \(\lnp_{it} * \lnpcdi_{it}\) | 0.272**  | (0.037)  |          |          |          |          |
| \(\lnfsa_{it}\)    | 0.36**   | (0.018)  | 0.380**  | (0.027)  | 0.393**  | (0.027)  |
| \(\lnpcdi_{it}\)   | -0.128   | (0.067)  | -0.209   | (0.072)  | -0.206   | (0.068)  |
| \(\lnp_{it}\)      | -0.004   | (0.048)  | -0.096   | (0.040)  | -0.067   | (0.041)  |
| Sargan Test         | 32.75**  | (0.71)   | 158.45   | (0.302)  | 147.76   | (0.398)  |
| AR(1) Test          | -3.15**  | (0.002)  | -3.09    | (0.002)  | -3.095   | (0.002)  |
| AR(2) Test          | -1.12**  | (0.263)  | -1.08    | (0.283)  | -1.05    | (0.294)  |
| Obs                 | 385      | 385      | 385      |          |          |          |

\(\star\): Note: the variable parenthesis is standard error, Sargan test and AR test parenthesis are corresponding P values. ***, **, and * mean significant at 1%, 5% and 10% levels, respectively.

The regression of model (1) results show that the goods in the issue of inventory has significant effects on the current inventory. Between the current inventory and inventory of the previous period have a positive correlations. House price and land supply have significant positive influence to inventory. So with the increase of house prices and land supply, inventory is also on the rise. Interest rates have a significant negative impact on inventory, with the increase of interest rate, the real estate enterprise financing costs, thereby reducing the supply of housing and inventory faster. Population density had no significant impact on the inventory. The model through the AR test and Sargan test shows that the selection of instrumental variables is reasonable and the model fitting effect is good. The regression of model (2) results show that after increasing prices squared, all variables impact on inventory were significantly. The coefficient of each variable symbol is consistent and have an order of magnitude with the model (1). House prices square coefficient is positive means that presents “U” shaped relationship between house prices and inventory, the price on the influence of the inventory may be nonlinear. The regression of Model (3) results show that after increasing the interaction item between housing price and per capita disposable income of urban residents, all variables have a significant impact on inventory, and the model not only has passed AR test and Sargan test but also has a good fitting effect. The regular interaction coefficient indicates that the increase of per capita disposable income can enhance the impact of housing price changes on inventory. The impact of house prices and land supply have significant positive influence to inventory, and economic growth has a positive regulating effect on the relationship between housing price and inventory.

2) Estimation of threshold effect of housing price on inventory

According to the regression results of the above model, there is not only a non-linear relationship between housing price and inventory, but also the per capita disposable income of urban residents can amplify the impact of housing price changes on inventory. Therefore, the variable of per capita disposable income of urban residents will be taken as the threshold variable, and the threshold regression model will be adopted to specifically test the non-linear relationship between housing price and inventory. The econometric model is set as follows:

\[
\lnfsa_{it} = \beta_0 + \beta_1 \lnp_{it} I(pcdi_{it} \leq \lambda_1) + \beta_2 \lnp_{it} I(pcdi_{it} < \lambda_1 \leq \lambda_2) + \beta_3 \lnp_{it} I(pcdi_{it} > \lambda_2) + \beta_4 C_{it} + \epsilon_{it} \tag{8}
\]
In equation (8), \( I(\cdot) \) is an index function. When the condition is satisfied, take 1; otherwise, take 0. \( \lambda \) is the threshold value; \( C \) is a series of factors that affect inventory (including population density, interest rate and land supply). Hansen's method was adopted in this paper to conduct the threshold of “grid search” according to the principle of minimum residual sum of squares, and the “self-sampling” method was used to simulate the likelihood ratio test statistics, and the P value was compared with its critical value (Hansen, 1999). The search order of threshold values is first single threshold, then double threshold and triple threshold. "Table IV" shows the corresponding threshold effect test results:

| Null hypothesis (H0) | F statistic | P-value | The inspection results | The number of Bootstrap | Threshold value |
|----------------------|-------------|---------|------------------------|-------------------------|-----------------|
| No threshold         | 52.19       | 0.0033  | Refuse                 | 300                     | 10.0969         |
| One threshold        | 26.49       | 0.0100  | Refuse                 | 300                     | 10.4072         |
| Two thresholds       | 8.45        | 0.5800  | accept                 | 300                     |                 |

"Table IV" shows that on the threshold regression of 35 large and medium cities nationwide, the F statistics obtained are respectively 52.19, 26.49 and 8.45 by the Bootstrap method (300). At the same time the P value is 0.0033, 0.0100 and 0.5800 respectively, which significantly declined “no threshold” and “there is a threshold method” but not rejected “there are two threshold values” assumption. Therefore, we can think that there are two threshold values, for the double threshold model. Specific threshold effect estimation results are shown in "Table V":

| lnsls | lnldp | lnlt | \( \beta_1 \) | \( \beta_2 \) | \( \beta_3 \) | R2(Within) | F Test | Number of samples |
|-------|-------|------|-------------|-------------|-------------|------------|--------|------------------|
| 0.630*** (0.057) | 0.176 (0.040) | -0.625*** (0.210) | 0.121 (0.090) | 0.161* (0.085) | 0.275*** (0.083) | 0.66 | 121.1 | 420 |

Note: variable parenthesis is standard error, "***", "**", and "*" mean significant at 1%, 5% and 10% levels, respectively.

As can be seen from "Table V", overall, the impact of housing price on inventory is positive, indicating that with the rise of housing price, inventory is increasing. By 35 large and medium cities of the country's threshold regression results can be seen that when the urban per capita disposable income is relatively low (\( \ln\text{pdci}; < 10.0969 \)), the relationship between prices and inventory was not significant; but with the increase of urban per capita disposable income, housing prices rose to the level of inventories is increasing, it is consistent with the regression of the model (3) results. It shows that with the development of economy, the per capita disposable income of urban residents is gradually increasing, and the positive effect of inventory increase caused by the rising housing price is increasing. The reason for such a result may be that although the per capita disposable income of urban residents is increasing, the increase rate of per capita disposable income of urban residents is far lower than the rate of housing price rise, and the increase of income is not enough to form an effective demand for house purchase, so that the inventory increases rapidly. In fact, it is also consistent with the recent high real estate inventory in most cities across the country. In addition, between the land supply and inventory have significant positive correlation shows that with the increase of land supply, inventory increased significantly; while the negative relationship between interest rate and stock, as interest rates decline, the stock has risen, this may be for real estate enterprises, loan interest rates will decrease the cost of capital of real estate enterprises and increase of housing supply, so indirectly led to the increase of inventory; while the population density of inventory no significant effect.

V. CONCLUSION

By building a nonlinear dynamic panel measurement model and using data of 35 large and medium-sized cities from 2005 to 2016, this paper first verified the existence of non-linear relationship between housing price and inventory, and then conducted threshold regression on the modified model to verify the threshold effect between housing price and inventory. Results show that the double threshold effect between house prices and inventory, when the urban per capita disposable income is low, the positive relationship between house prices and the stock is not significant; but with the increase of urban per capita disposable income, housing prices for positive influence on strengthening inventory, illustrate urban per capita disposable income rose far lagged behind the house prices rise. It makes residents housing demand is insufficient and effectively making real estate inventory increase quickly.

At present, the real estate destocking control policy at the national level has come to an end. The following points should be paid attention to in the future: first, the real estate market. First, lower house prices. In cities with large inventory, the government should focus on suppressing the sharp rise in housing prices, guiding housing prices to return to a reasonable range in the long run, increasing residents' income level, and effectively improving the imbalance between supply and demand in the housing market. Second, reducing financial leverage in the real estate market. At present our country is in a critical period of financial deleveraging, in order to reduce the financial risk of real estate market, not only to reduce financial institutions from the source to the various real estate development...
enterprise land capital credit support, strict controls of the real estate credit eventually flow, but also by raising the lending rates to reduce the real estate enterprise financing lever. Third, increasing land supply. With the implementation of government policies on a series of control housing prices in recent years, some recent first-tier cities inventory to effect is obvious, therefore, as for lower inventory city, the government should pay more attention to by increasing land supply to improve housing security, then foster and develop the housing rental market, so as to ensure the reasonable housing needs of low-income residents.

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