Association between low dietary zinc and hyperuricaemia in middle-aged and older males in China: a cross-sectional study

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ABSTRACT

Objective: To examine the associations between dietary zinc intake and hyperuricaemia.

Design: Cross-sectional study.

Setting: This study was conducted in a health examination centre of China.

Participants: A total of 5168 middle-aged and older participants (aged 40 years or above) (2697 men and 2471 women) were included.

Outcome measures: Dietary zinc intake was assessed using a validated semiquantitative food frequency questionnaire. Hyperuricaemia was defined as uric acid $\geq 416$ μmol/L for males and $\geq 360$ μmol/L for females.

Results: For males, the prevalence of hyperuricaemia was 22.9%. After adjusting for age, body mass index (BMI) and energy intake, the ORs were 0.68 (95% CI 0.45 to 0.92) in the second quintile, 0.63 (95% CI 0.45 to 0.89) in the third quintile, 0.68 (95% CI 0.46 to 1.00) in the fourth quintile and 0.55 (95% CI 0.35 to 0.87) in the fifth quintile comparing the lowest quintile of Zn intake, respectively (p for trend=0.03). In the multivariable adjusted model, the relative odds of hyperuricaemia were significantly decreased by 0.71 times in the second quintile of zinc intake (OR 0.71, 95% CI 0.52 to 0.98), 0.64 times in the third quintile (OR 0.65, 95% CI 0.44 to 0.94) and 0.55 times in the fifth quintile (OR 0.56, 95% CI 0.32 to 0.97) compared with those in the lowest quintile, and p for trend was 0.064. For females, the prevalence of hyperuricaemia was 10.0%, and unadjusted, minimally adjusted as well as multivariable adjusted ORs all suggested no significant association between dietary zinc intake and hyperuricaemia.

Conclusions: The findings of this cross-sectional study indicated that dietary zinc intake was inversely associated with hyperuricaemia in middle-aged and older males, but not in females. The association was significant after considering the influence of age, BMI and energy intake, and after that, minimum adjustment remained independent of further confounding factors such as vitamin C intake, alcohol drinking status and nutrient supplementation.

INTRODUCTION

Hyperuricaemia (HU) occurs when serum uric acid, the final product of purine metabolism, exceeds the normal level. Previous studies have demonstrated close associations between HU and increased risks of gout as well as other medical problems such as type 2 diabetes, metabolic syndrome and cardiovascular disease. Over the past several decades, the prevalence of HU has increased worldwide. In recent epidemiological studies, the reported prevalence of HU ranged from 8.9% to 24.4% in different populations. Therefore, HU has emerged as a major health problem and is currently gaining more attention. However, the pathophysiology of HU has not been fully understood.

Zinc (Zn) is an essential trace element that plays a wide range of biological roles in catalytic, structural and regulatory functions. Acting as an antioxidant and anti-inflammatory agent in humans, Zn may be associated with the level of serum uric acid, which may function as a pro-oxidant and proinflammatory factor under certain circumstances. Umeki et al. revealed that oral Zn administration could normalise serum uric acid in a patient with Wilson’s
disease by improving liver dysfunction. In several animal studies, Zn supplementation or administration of Zn complex also showed normalisation of elevated serum uric acid levels in rats with aspirin-related damage, experimental diabetes or intestinal injury.\textsuperscript{14–16} Besides, a longitudinal study by Navarro-Alarcon \textit{et al}\textsuperscript{17} found that uric acid was negatively linearly related to serum Zn in haemodialysis patients. In view of these factors, Zn deficiency may be related to HU to a certain extent. However, it should be noted that an animal study conducted by Bruno \textit{et al}\textsuperscript{18} observed that plasma uric acid concentrations were significantly decreased among rats with zinc deficiency and a vitro study conducted by Bataille \textit{et al}\textsuperscript{19} demonstrated that zinc-induced cellular stress could decrease avian renal proximal tubule uric acid secretion and may therefore increase plasma uric acid concentration. Therefore, the effect of Zn on serum uric acid level still remains controversial. To the best of our knowledge, there has been no study conducted directly examining the association between dietary Zn intake and the prevalence of HU. The purpose of this cross-sectional study was to explore the aforementioned association based on the following hypothesis: dietary Zn intake is negatively associated with the prevalence of HU.

\textbf{MATERIALS AND METHODS}

\textbf{Study population}

This cross-sectional study was conducted in the Department of Health Examination Center, Xiangya Hospital, Central South University, in Changsha, Hunan Province, China. We obtained approval for this study from the ethics committee at Xiangya Hospital, Central South University. We also obtained written informed consent from the participants in our study. The study design has been published previously.\textsuperscript{20} Routine health check-ups are very common in China, because the Chinese government encourages people to take periodic medical examinations. Registered nurses interviewed all participants during the examination, using a standard questionnaire, with the purpose of collecting information on demographic characteristics and health-related habits. Participants were selected according to the following inclusion criteria: (1) they were aged 40 years or above; (2) they were undergoing serum uric acid measurement; (3) they had completed the semiquantitative food frequency questionnaire (FFQ) on the average amount of food consumed on each occasion: less than one to three times per week, one to three times per week, four to five times per week, once per day, twice per day, or three times and above per day) they consumed each food item in the past year. There were six options for the average frequency of physical activity (never, one to two times per week, three to four times per week, five times and above per week) and (4) all the participant’s basic characteristics, including age, gender, body mass index (BMI), smoking status, alcohol drinking status, etc, were available. In the beginning, this cross-sectional study included 10,387 participants who were undergoing routine check-ups, including serum uric acid measurement, at the Department of Health Examination Center, Xiangya Hospital, Central South University, in Changsha, Hunan Province, China, from October 2013 to July 2014. At the time, 9420 individuals were aged over 40 years, and 9402 had basic characteristics, such as BMI, available. Eventually, 5168 participants completed the FFQ, and the overall response rate of the survey was 54.9%. There was no significant difference between participants who completed and those who did not complete the FFQ in terms of the prevalence of HU (\textit{p}=0.21) and in terms of characteristics, including age (\textit{p}=0.12) and BMI (\textit{p}=0.19).

\textbf{Assessment of HU}

All blood samples were drawn after a 12 h overnight fast and were kept at 4°C until analysis. Uric acid was detected on a Beckman Coulter AU 5800 system (Beckman Coulter Inc, Brea, California, USA). HU was defined by uric acid \(\geq 416\, \mu\text{mol/L} \) in males and \(\geq 360\, \mu\text{mol/L} \) in females.

\textbf{Assessment of dietary and non-diary exposures}

Dietary intake was evaluated using a semi-quantitative food frequency questionnaire (SFFQ) that was specially designed for the population in Hunan province of China. It has been used and validated in several previously published studies.\textsuperscript{21, 22} This SFFQ contained 63 food items that are regularly consumed in Hunan province. Participants were requested to answer how frequently (never, once per month, two to three times per month, one to three times per week, four to five times per week, once per day, twice per day, or three times and above per day) they consumed each food item in the past year. There were six options for the average amount of food consumed on each occasion: less than 100, 100–200, 201–300, 301–400, 401–500, and over 500 g. Colour pictures showing food samples with labelled weights were given to participants as a reference. The SFFQ was either self-administered or completed through interviews by professional researchers. The validity of the SFFQ was tested through comparing with the 24 h dietary recall method in 91 participants, where the tested samples were randomly selected from the same study population. The correlation coefficient between the SFFQ and the 24 h recall was 0.34 for the measurement of zinc intake. The strength of the correlation coefficient was also close to other similar published studies.\textsuperscript{23–25} The Chinese Food Composition Table was referenced to calculate the individual composition of macronutrients and micronutrients of the included foods.\textsuperscript{26} This SFFQ has been validated and used in a previously published study.\textsuperscript{27} The weight and height of each participant were respectively measured to calculate the BMI. Participants were also asked about their average frequency of physical activity (never, one to two times per week, three to four times per week, five times and above per week) and average duration of physical activity (less than half an hour, half an hour to 1, 1–2, more than 2 h). Smoking and alcohol drinking status was asked about face to face.

Blood fasting glucose, high density lipoprotein cholesterol (HDL-cholesterol), low density lipoprotein
cholesterol (LDL-cholesterol) and triglyceride (TG), were also detected on the Beckman Coulter AU 5800 system (Beckman Coulter Inc, Brea, California, USA). Blood pressure was measured using an electronic sphygmomanometer. Participants with fasting glucose of ≥7.0 mmol/L or currently undergoing drug treatment for blood glucose control were regarded as diabetes patients, and participants with systolic blood pressure of ≥140 mm Hg or diastolic blood pressure of ≥90 mm Hg, or currently using antihypertensive medication, were regarded as hypertension patients.

**Statistical analysis**

The continuous data are expressed as mean±SD (normally distributed data) or median and IQR (skewed distributed data), and the category data are expressed in percentage. Zn intake was classified into five categories based on the quintile distribution: ≤16.23, 16.24–19.07, 19.08–21.95, 21.96–22.55 and ≥22.56 mg/day in males; and ≤13.63, 13.64–15.98, 15.99–18.53, 18.54–21.99 and ≥22.00 mg/day in females. The ORs with 95% CIs for the association between HU and dietary Zn were calculated for each quintile of Zn intake, respectively, and the quintile with the lowest value was regarded as the reference category. In order to calculate the adjusted OR values of each quintile of Zn intake, a minimally adjusted model and a multivariable adjusted model were adopted in the logistic analyses for males and females, respectively. Covariant variables in the minimally adjusted model included age, BMI and total energy intake. Covariant variables in the multivariable model included age, BMI, educational level, activity level, total energy intake, fibre intake, vitamin C intake, vitamin E intake, smoking status, alcohol drinking status, nutrient supplementation, diabetes, hypertension, HDL-cholesterol, LDL-cholesterol and TG. Covariates were chosen based on some published studies. The female subgroup was also adjusted for menopausal status. Tests for linear trends were conducted based on logistic regression, using a median variable of Zn level in each category. Model fitness was assessed by a $\chi^2$ test and using a maximum likelihood method (p=0.000). We also examined the multicollinearity among the covariates by correlation coefficient matrix, no significant multicollinearity (r>0.8) existed. Sensitivity analysis was conducted by excluding participants who had mineral (except calcium) supplementation. We also conducted a subgroup analysis by stratifying the data of participants with or without hypertension or diabetes. All data analyses were performed using SPSS V17.0; a p value equal to or less than 0.05 was considered to be statistically significant.

**RESULTS**

The basic characteristics of the study population in men and women are listed in table 1. A total of 5168 participants (2697 men and 2471 women) were included in the present study. The overall prevalence of HU among the participants of this cross-sectional study (aged 40–85 years, average age=50.1±7.6) was 16.7%. Significant differences between males and females were observed for the prevalence of HU, age, BMI, smoking status, alcohol drinking status, energy intake, fibre intake, vitamin intake, nutrient supplementation, education level, diabetes, hypertension, HDL-cholesterol and TG.

Table 2 presents the associations between HU and dietary Zn intake in males and females. For males, the prevalence of HU was 22.9% in this study. The unadjusted ORs suggested that no significant association existed between Zn intake and HU. However, after adjusting for age, BMI and energy intake, the results suggested a strong inverse association between Zn intake and HU. The ORs were 0.68 (95% CI 0.45 to 0.92) in the second quintile, 0.63 (95% CI 0.45 to 0.89) in the third quintile, 0.68 (95% CI 0.46 to 1.00) in the fourth quintile and 0.55 (95% CI 0.35 to 0.87) in the fifth quintile comparing the lowest quintile of Zn intake, respectively (p for trend=0.05). In the multivariable adjusted model, the relative odds of HU were significantly decreased by 0.71 times in the second quintile of Zn intake (OR 0.71, 95% CI 0.52 to 0.98, p=0.039), 0.65 times in the third quintile (OR 0.65, 95% CI 0.44 to 0.94, p=0.023) and 0.56 times in the fifth quintile (OR 0.56, 95% CI 0.32 to 0.97, p=0.038) compared with those in the lowest quintile, and p for trend was 0.064. For females, the prevalence of HU was 10.0% in this study. Unadjusted, minimally adjusted and multivariable adjusted ORs all suggested that no significant association was observed between dietary zinc intake and HU. The multivariable-adjusted ORs (95% CI) of HU across the five quintiles of Zn intake were 1.70 (95% CI 1.07 to 2.71), 1.19 (95% CI 0.67 to 2.11), 1.05 (95% CI 0.52 to 2.11) and 1.38 (95% CI 0.59 to 3.25) respectively, while p for trend was 0.582. After excluding participants with mineral (except calcium) supplementation, the outcomes did not change significantly in males and females.

The results of subgroup analysis are listed in table 3. In males without hypertension or diabetes, the inverse association between Zn intake and HU still existed (p for trend=0.03). But there was no association between Zn intake and HU in males with hypertension or diabetes (p for trend=0.57). The outcomes suggest that the existence of hypertension or diabetes may weaken the association between Zn intake and HU. Outcomes from both female subgroups showed that there were no significant associations between Zn intake and HU.

**DISCUSSION**

This cross-sectional study revealed that Zn intake was inversely associated with the prevalence of HU in male participants. The association was significant after considering the influence of age, BMI and energy intake, and, after that, minimum adjustment remained independent of further confounding factors such as vitamin C intake, alcohol drinking status and nutrient supplementation.
Besides, the inverse association noted between Zn intake and HU was only observed in males without hypertension or diabetes but not in those suffering from these two diseases. In females, however, no significant association was observed between dietary zinc intake and HU.

High prevalence of HU in the Chinese population may be related to recent dietary changes. Previous published data have shown that an increased intake of meat and seafood, or alcohol consumption, significantly increased the risk of developing HU. On the other hand, randomised controlled trials revealed that some dietary interventions, such as dairy intake and vitamin C supplementation, might be protective against the development of HU. Therefore, dietary modification is often recommended for patients with HU.

The findings of this cross-sectional study demonstrated that Zn intake was inversely associated with the prevalence of HU in males, indicating a potentially beneficial role of Zn intake in the development of HU in males. Because of its relatively high concentration compared with other antioxidants and its ability to scavenge oxygen radicals, uric acid is often, at physiological concentrations, considered a major hydrophilic antioxidant in human plasma. One of the main sites where the antioxidant property of uric acid has been demonstrated is the central nervous system, as neurons are remarkably susceptible to oxidative damage. Previous studies focusing on neurological disease such as multiple sclerosis, Parkinson’s disease and ischaemic stroke have suggested some potential protective role in raising serum uric acid. However, some studies have reported that HU was associated with an increased risk of some diseases related to oxidative stress, such as metabolic syndrome, cardiovascular disease and obesity. The possible explanation may be that uric acid could stimulate oxidative stress through the renin-angiotensin system (RAS) in vascular endothelial cells as well as vascular smooth muscle cells, and increase nicotinamide adenine dinucleotide phosphate oxidase-derived reactive oxygen species production. Taken together, these findings provide evidence that uric acid may function as a pro-oxidant in HU, despite its antioxidant properties under physiological conditions. The fact that Zn, the most abundant trace intracellular element, has the ability to retard the oxidative process, has been recognised for many years. Taken together, these findings indicated that Zn intake and HU were inversely associated, and the antioxidant properties of Zn might be protective against the development of HU.

Table 1 Characteristics among 5168 participants in the male and female population

| Characteristics | Male population (n=2697) | Female population (n=2471) | p Value |
|-----------------|--------------------------|---------------------------|---------|
| Median Zn intake (mg/d) | 20.54 | 17.20 | 0.00 |
| Hyperuricaemia (%) | 22.9 | 10.0 | 0.00 |
| Age (years) | 52.7±7.48, 51.04 (47.12–57.70) | 53.49±7.66, 51.70 (47.85–58.70) | 0.00 |
| Age range (years) | 40–85 | 40–84 | 0.00 |
| BMI (kg/m²) | 25.12±3.12, 25.09 (23.05–27.13) | 23.77±3.09, 23.64 (21.65–25.61) | 0.00 |
| Smoking (%) | 40.7 | 2.5 | 0.00 |
| Alcohol drinking (%) | 54.8 | 14.9 | 0.00 |
| Activity level (h/w) | 2.18±3.50, 0 (0–4.5) | 2.38±3.58, 0 (0–4.5) | 0.20 |
| Mean total energy intake (kcal/day) | 1814.78±827.38, 1689.96 (1325.94–2093.06) | 1468.61±683.40, 1358.03 (1057.54–1722.57) | 0.00 |
| Mean fibre intake (g/day) | 20.14±14.94, 16.68 (9.99–29.00) | 16.48±12.89, 12.65 (7.58–21.84) | 0.00 |
| Mean vitamin C intake (mg/day) | 116.42±72.38, 108.84 (72.38–143.31) | 119.50±75.41, 109.37 (71.73–144.36) | 0.55 |
| Mean vitamin E intake (mg/day) | 30.91±14.06, 28.61 (21.58–37.25) | 28.44±14.63, 25.34 (19.21–34.67) | 0.00 |
| Nutrient supplementation (%) | 27.1 | 43.5 | 0.00 |
| Calcium | 20.3 | 35.3 | 0.00 |
| Vitamins | 9.0 | 15.0 | 0.00 |
| Minerals (except calcium) | 0.7 | 0.8 | 0.66 |
| High school background or above (%) | 55.2 | 38.2 | 0.00 |
| Diabetes (%) | 11.9 | 7.3 | 0.00 |
| Hypertension (%) | 35.4 | 30.3 | 0.00 |
| HDL-cholesterol (mg/dL) | 1.41±0.35, 1.36 (1.17–1.59) | 1.64±0.38, 1.60 (1.36–1.88) | 0.00 |
| LDL-cholesterol (mg/dL) | 2.94±0.92, 2.91 (2.34–3.51) | 2.99±0.93 2.95 (2.37–3.56) | 0.17 |
| Triglyceride (mg/dL) | 2.17±2.06, 1.62 (1.11–2.40) | 1.58±1.34, 1.25 (0.91–1.84) | 0.00 |

Continuous data are showed as mean±SD and median (IQR), category data are shown as percentage. BMI, body mass index; HDL, high-density lipoprotein; LDL, low density lipoprotein; Zn, zinc.
### Table 2  Relations between dietary Zn intake and hyperuricaemia (n=5168)

| Quintiles of Zn intake | 1 (lowest) | 2 | 3 | 4 | 5 (highest) | p for trend |
|------------------------|------------|---|---|---|-------------|------------|
| **Males (n=2697)**     |            |   |   |   |             |            |
| Median Zn intake (mg/d)| 14.06      | 17.68 | 20.54 | 23.51 | 28.42 | –          |
| Participants (n)       | 539        | 540 | 539 | 540 | 539 | –          |
| Unadjusted ORs         | 1.00 (Reference) | 0.85 (0.64 to 1.13) | 0.86 (0.65 to 1.15) | 1.06 (0.80 to 1.40) | 1.00 (0.76 to 1.32) | 0.527 |
| Minimally adjusted ORs*| 1.00 (Reference) | 0.68 (0.45 to 0.92) | 0.63 (0.45 to 0.89) | 0.68 (0.46 to 1.00) | 0.55 (0.35 to 0.87) | 0.030 |
| Multivariable adjusted ORs† | 1.00 (Reference) | 0.71 (0.52 to 0.98) | 0.65 (0.44 to 0.94) | 0.71 (0.45 to 1.11) | 0.56 (0.32 to 0.97) | 0.064 |
| **Females (n=2471)**   |            |   |   |   |             |            |
| Median Zn intake (mg/day)| 11.82 | 14.91 | 17.21 | 20.00 | 24.97 | –          |
| Participants (n)       | 494 | 494 | 496 | 494 | 493 | –          |
| Unadjusted ORs         | 1.00 (Reference) | 1.52 (1.00 to 2.30) | 1.07 (0.69 to 1.67) | 1.03 (0.66 to 1.60) | 1.38 (0.91 to 2.10) | 0.478 |
| Minimally adjusted ORs*| 1.00 (Reference) | 1.61 (1.04 to 2.50) | 1.19 (0.70 to 2.02) | 1.13 (0.59 to 2.13) | 1.53 (0.71 to 3.31) | 0.356 |
| Multivariable adjusted ORs† | 1.00 (Reference) | 1.70 (1.07 to 2.71) | 1.19 (0.67 to 2.11) | 1.05 (0.52 to 2.11) | 1.38 (0.59 to 3.25) | 0.582 |

Data are adjusted ORs (95% CIs), unless otherwise indicated.

*Minimally adjusted model was adjusted for age, BMI and total energy intake.

†Multivariable adjusted model was adjusted for age, BMI, educational level, activity level, total energy intake, fibre intake, vitamin C intake, vitamin E intake, smoking status, alcohol drinking status, nutrient supplementation, diabetes, hypertension, HDL-cholesterol, LDL-cholesterol and triglyceride. The female subgroup was also adjusted for menopausal status.

BMI, body mass index; HU, hyperuricaemia; n, number; Zn, zinc.

### Table 3  Subgroup analysis by stratifying the data of participants with or without hypertension or diabetes

| Quintiles of Zn intake | 1 (lowest) | 2 | 3 | 4 | 5 (highest) | p for trend |
|------------------------|------------|---|---|---|-------------|------------|
| **Males without hypertension or diabetes (n=1587)** |            |   |   |   |             |            |
| Multivariable adjusted ORs* | 1.00 (Reference) | 0.68 (0.45 to 1.05) | 0.52 (0.31 to 0.86) | 0.48 (0.26 to 0.90) | 0.41 (0.19 to 0.90) | 0.030 |
| **Males with hypertension or diabetes (n=1110)** |            |   |   |   |             |            |
| Multivariable adjusted ORs* | 1.00 (Reference) | 0.69 (0.42 to 1.12) | 0.79 (0.45 to 1.39) | 1.02 (0.53 to 1.97) | 0.72 (0.32 to 1.63) | 0.572 |
| **Females without hypertension or diabetes (n=1645)** |            |   |   |   |             |            |
| Multivariable adjusted ORs* | 1.00 (Reference) | 1.49 (0.73 to 3.02) | 1.21 (0.52 to 2.82) | 0.91 (0.32 to 2.57) | 1.20 (0.34 to 4.17) | 0.847 |
| **Females with hypertension or diabetes (n=826)** |            |   |   |   |             |            |
| Multivariable adjusted ORs* | 1.00 (Reference) | 1.79 (0.97 to 3.32) | 1.14 (0.53 to 2.48) | 1.07 (0.42 to 2.73) | 1.43 (0.44 to 4.66) | 0.671 |

Data are adjusted ORs (95% CIs), unless otherwise indicated.

*The multivariable adjusted model was adjusted for age, BMI, educational level, activity level, total energy intake, fibre intake, vitamin C intake, vitamin E intake, smoking status, alcohol drinking status, nutrient supplementation, HDL-cholesterol, LDL-cholesterol and triglycerides. The female subgroup was also adjusted for menopausal status.

BMI, body mass index; HDL, high-density lipoprotein; LDL, low density lipoprotein n, number; Zn, zinc.
are needed to explore it. When a subgroup analysis stratified by males with hypertension or diabetes was performed, only those without hypertension or diabetes showed an inverse association between dietary Zn intake and HU. Several studies have showed that hypertension may be an independent risk factors for HU.7 52 Similarly, A recent review by Li et al39 concluded that insulin resistance and HU shared a bidirectional causal effect. This collection of evidence may provide an explanation for how these two diseases could attenuate the inverse association between dietary Zn intake and HU.

An interesting finding of this study is that the inverse effect of dietary Zn intake on HU is only valid for males, and not for females. One hypothesis is that the fractional zinc absorption in men is relatively greater compared to that in women. It has been proposed that men are more easily affected by Zn deficiency, a public health problem with an estimated prevalence of 17.3% among the world’s population.50–52 Furthermore, Zn deficiency, as reported by previous studies, is associated with increased fractional Zn absorption.53 Taken together, fractional Zn absorption in men may be higher than in women, and therefore contribute to its protective effects against HU in men. Another potential explanation for the difference in effect by gender is that the high prevalence of zinc deficiency in males may make them more susceptible to the consequences of zinc deficiency. However, the current study did not confirm these speculations and further studies are needed to clarify the gender difference.

The present study has several strengths. First, this is the only study, to date, conducted on a large sample (5168 participants), to explore the association between dietary Zn intake and HU. Second, the multivariable model was adjusted for a considerable number of potentially confounding factors, such as diabetes, hypertension and blood lipid levels, which greatly improved the reliability of the results. Third, this study adopted FFQ, an effective measurement method for dietary micronutrient intake,27 54 55 which could estimate annual dietary intake. It might represent the average level of micronutrients more effectively than the measurement of instant serum level can.

Limitations of the present study should also be acknowledged. First, as is well recognised, flesh foods are the major sources of purine as well as bioavailable dietary zinc, and uric acid is the end product of purine metabolism. Accordingly, the inverse association observed between dietary Zn intake and HU may be due to the indirect effect of purine intake. Unfortunately, the data of the present study were not available for calculating purine intake. Second, this study provided only cross-sectional data, which precluded causal correlations, and, thus, further prospective longitudinal studies and intervention trials should be undertaken to establish a causal association between dietary Zn intake and HU. Since no previous research investigated the association between dietary Zn intake and HU, the value of this study should not be blotted out due to its cross-sectional nature. Third, another limitation lies in the FFQ method itself, which is a relatively less accurate method for instantly reflecting the level of micronutrients when compared with blood concentration measurement. However, blood level may not fully reflect nutritional status,56 and it is also difficult to reflect the accurate circulating level after a certain period of time. Finally, further studies are needed to explore the mechanism of this correlation.

CONCLUSIONS

The findings of this cross-sectional study indicated that dietary Zn intake was inversely associated with HU in middle-aged and older males, but not in females. The association was significant after considering the influence of age, BMI and energy intake, and after that, minimum adjustment remained independent of further confounding factors such as vitamin C intake, alcohol drinking status and nutrient supplementation. As the study is cross sectional, its ability to prove causation is limited. Hence further studies based on rigorous designs should be recommended.

Contributors G-hL, D-xX, Y-IX and CZ collected the data. G-hL wrote the manuscript. D-xX, Y-IX and CZ helped in writing the manuscript. D-xX and Y-IX conducted statistical analysis and revised the manuscript. JW, TY, HL Y-IW, S-gG and Y-sL reviewed the manuscript. All authors contributed to the design of the study.

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Competing interests None declared.

Patient consent Obtained.

Ethics approval The study was approved by the ethics committee at Xiangya Hospital, Central South University.

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