Increased expression of Ki-67 is a poor prognostic marker for colorectal cancer patients: a meta analysis

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Abstract

Background: The prognostic value of Ki-67 expression in colorectal cancer patients was controversial. Therefore, this meta analysis was conducted to ascertain the prognostic value of Ki-67 expression in colorectal cancer patients.

Methods: The electronic databases, including EMBASE, PubMed, Cochrane Library and Web of Knowledge database, were searched from January 1970 to July 2017. The pooled hazard ratios and 95% confidence intervals were calculated to evaluate the prognostic value of Ki-67 expression for colorectal cancer patients.

Results: Totally 34 eligible studies and 6180 colorectal cancer patients were included in the present meta analysis. The pooled hazard ratios were 1.54(95% CI 1.17–2.02, P = 0.005) for overall survival and 1.43(1.12–1.83, P = 0.008) for disease free survival in univariate analysis. After adjustment of other prognostic factors, the pooled HR was 1.50(95% CI 1.02–2.22, P = 0.03) for overall survival in multivariate analysis.

Conclusion: The present meta analysis demonstrated that high Ki-67 expression is significantly correlated with poor overall survival and disease free survival, indicating that high Ki-67 expression may serve as a valuable predictive method for poor prognosis of colorectal cancer patients.

Introduction

Colorectal carcinoma (CRC) is the fourth most common cause of death induced by carcinoma, accounting for 600,000 deaths per year [1]. Some clinical factors, such as tumor stage, tumor necrosis, vascular invasion and differentiation, have been reported to be associated with prognosis of CRC patients [2, 3]. However, these factors were not sufficient to accurately make risk stratification for clinical prognosis. In clinical practice, identification of patients with high risk of poor prognosis would help to optimize treatment and to improve clinical prognosis. Therefore, there is an urgent need to find reliable prognostic factors for prediction of clinical prognosis in CRC patients.

Ki-67 antigen expresses throughout the cell cycle in proliferating cells, except for quiescent (G0) cells [4]. The correlation between Ki-67 expression and prognosis of CRC patients were still contradictory in various studies [5–38]. Several studies showed that high Ki-67 expression was related with poor overall survival (OS) [26–30], whereas several studies reported that high Ki-67 expression was correlated with favorable OS [8, 15, 16]. Therefore, we performed this meta analysis to determine the prognostic value of Ki-67 expression in CRC patients.

Materials and methods

Literature search strategy

Four electronic databases, including EMBASE, PubMed, Web of Knowledge and Cochrane Library, were searched from January 1990 to July 2017 for eligible studies. We performed literature search by combined text word and MeSH strategy for PubMed with the following search strategy: (“Ki67” or “Ki-67” or “MIB-1” or “MIB1”) and (“carcinoma” or “neoplasm” or “tumor” or “cancer” or “malignancy”) and (“colorectal” or “rectal” or “colon”) and (“prognosis” or “prognostic” or “survival” or “outcome” or “mortality”). The strategies for EMBASE and other databases were similar but were adapted according to the guideline of the database. Meanwhile, expanded search of hyponym was performed. In the retrieval
process, we made a manual search using the references of the including articles to supplement eligible studies. We contacted the corresponding author to get necessary study data if necessary. The search was restricted to human studies and there was no restriction in terms of language or publication time. All information investigation and collection were conducted according to the principles of Declaration of Helsinki.

Criteria for inclusion and exclusion
The inclusion criteria were defined as follows: (1) proven pathological diagnosis of CRC in humans; (2) assessment of Ki-67 expression by using immunohistochemistry (IHC) method; (3) enough survival information such as hazard ratio (HR) and 95% confidence interval (CI). Studies not directly providing hazard ratio and 95% confidence interval were included if survival information were available from survival curves or tables. Articles published in Chinese were included in the present meta analysis as English literature. For multiple studies from the same study population, only the most recently published study was included in the present analysis.

The following studies were excluded: (1) non-human experiments; (2) case reports, reviews, letters and conference abstracts without survival information; (3) lack of the necessary survival information.

Quality assessment
Two reviewers (Zhi-Qiao Zhang and Zhao-Wen Luo) independently assessed the quality of the studies included in the present meta analysis using Newcastle-Ottawa Quality Assessment Scale (NOS). The NOS comprises assessments of patient selection, study comparability, follow-up and outcomes of interest. The total scores were used to evaluate the quality of the study. Disagreements between two reviewers (Zhi-Qiao Zhang and Zhao-Wen Luo) were resolved by consultation with a third reviewer (Ming-Gu Zhu).

Data extraction
Two investigators (Zhi-Qiao Zhang and Zhao-Wen Luo) independently extracted the survival information of the original studies: surname of the first author, country, study sample size, publication year, disease stage, test method of Ki-67 expression, clinical parameters, and survival information (HRs and 95%CIs). All study information were extracted and recorded by using a standardized form. All included studies were coded as the surname of the first author + publish year in the standardized form. Study authors were contacted to get necessary data if necessary. Disagreements between two investigators were resolved by discussion with a third investigator (Ming-Gu Zhu).

Statistical analysis
The statistical analysis was performed according to the suggestions advised by the Meta analysis of Observational. Studies in Epidemiology group(MOOSE) [39]. The hazard ratios (HR) and 95% CI were used to summary the final effects of survival outcome. We pooled HR and 95% CI if the survival data were reported in the original articles. While HR and 95% CI were not directly reported in the articles, survival data was extracted from Kaplan-Meier curve and used to estimate HR. The heterogeneity among various studies was measured by the Q and I² tests. A probability value of I² value≥30% and P value< 0.1 indicated the existence of significant heterogeneity. A random effect model (DerSimonian and Laird method) or fixed effect model(Mantel-Haenszel method) was used according to the results of heterogeneity analyses. Meta-regression analyses (Reml method) and subgroup analyses were performed to explore the sources of heterogeneity. Funnel plot, Begg’s test, Egger’s test, and trim and fill method were performed to assess publication bias. The difference was considered to be statistically significant if P value < 0.05. The statistical analyses were performed by STATA version 12.0 software (Stata Corporation, College Station, Texas, USA).

Results
Search results
The initial search returned a total of 410 articles (with 204 duplicate articles). After screening the abstracts, 159 irrelevant articles were excluded according to the criteria for inclusion and exclusion. Reviewers identified 47 potential studies for full-text review and 13 articles were eliminated due to inadequate data for meta analysis. Finally, a total of 34 studies were included in the present meta analysis [5–38]. The details of screening process were shown in Fig. 1. Quality assessment of 34 eligible studies were performed by using the Newcastle-Ottawa Scale (NOS).

Study selection and characteristics
The characteristic of the 34 included studies were summarized in Table 1. The publication time ranged from 1997 to 2015. The patient number of 34 studies ranged from 30 to 867, with a mean sample size of 182. The mean length of follow-up period ranged from 23 to 130 months. The NOS scores for quality of 34 studies in this meta analysis varied from 6 to 7, with a mean value of 6.5. Ki-67 expression were measured in surgical tumor tissues.

Prognostic value of high Ki-67 expression
Finally 34 studies and 6180 CRC patients were collected and analyzed for prognostic value of Ki-67 expression. The pooled HR was 1.54(95% CI 1.17–2.02, P = 0.005)
for overall survival (OS) in univariate analysis (Fig. 2). After adjustment of other prognostic factors, the pooled HR was 1.50 (95% CI 1.02–2.22, \( P = 0.03 \)) for OS in multivariate analysis (Fig. 3). The pooled HR was 1.43 (95% CI 1.12–1.83, \( P = 0.008 \)) for disease survival (DFS) in univariate analysis (Fig. 4).

**Publication bias**

For OS in univariate analysis, the funnel plot of Ki-67 expression showed obvious asymmetry (Fig. 5a), indicating that the pooled results might be influenced by the publication bias. The Begg’s test and Egger’s test were further performed to explore publication bias in the present meta analysis. There was significant publication bias according to Egger’s test (\( P = 0.008 \)) whereas Begg’s test did not show significant publication bias (\( P = 0.629 \)).

For OS in multivariate analysis, the funnel plot of Ki-67 expression did not show obvious asymmetry (Fig. 5b). There was no significant publication bias according to Egger’s test (\( P = 0.378 \)) and Begg’s test (\( P = 1.0 \)).

For DFS in univariate analysis, the funnel plot of Ki-67 expression showed obvious asymmetry (Fig. 5c). There was significant publication bias according to Egger’s test (\( P = 0.032 \)) whereas Begg’s test did not show significant publication bias in the present study (\( P = 0.246 \)).
| Author/Country/Year | Detection method | Quantitative method | Antibody type | Fixed method | Positivea Rate(%) (Range) | Median age (Range) | History | Therapy Type/Stage | Cut off Value | Patient Number |
|---------------------|------------------|---------------------|---------------|-------------|--------------------------|-------------------|---------|------------------|--------------|---------------|
| Meyer [5] Germany 2009 | microscope | semi-quantitative | mouse monoclonal | formalin | 88.8% 64(28–91) | CC II-IV | S + C | NR | 392 |
| Garrity [6] Canada 2004 | NR | quantitative | NR | formalin | 62.4% NR | CRC II-III | S + C | 27.0% | 366 |
| Fodor [7] America 2012 | Tissue microarrays | quantitative | mouse monoclonal | formalin | 50.2% 71.5(18–97) | CRC I-IV | S + C | 60.0% | 867 |
| Salmiminen [8] Finland 2005 | microscope | quantitative | mouse monoclonal | formalin | 45.2% 66(36–86) | RC I-IV | S + R | 40.0% | 146 |
| Zhang [9] China 2015 | Tissue microarrays | quantitative | NR | formalin | 47.9% 72(37–93) | CRC I-IV | S + R | 40.0% | 73 |
| Salminen [8] Finland 2005 | microscope | quantitative | mouse monoclonal | formalin | 45.2% 66(36–86) | CRC I-IV | S + R | 40.0% | 73 |
| Zhang [9] China 2015 | Tissue microarrays | quantitative | NR | formalin | 47.9% 72(37–93) | CRC I-IV | S + R | 40.0% | 73 |
| Li [11] China 2010 | Tissue microarrays | semi-quantitative | monoclonal | formalin | 78.8% 65(22–95) | CRC I-IV | S | 10.0% | 203 |
| Li [12] China 2011 | Tissue microarrays | semi-quantitative | monoclonal | formalin | 78.8% 65(22–95) | CRC I-IV | S | 10.0% | 203 |
| Dawson [13] Switzerland 2014 | Tissue microarrays | quantitative | monoclonal | formalin | 53.3% 66(29–79) | CRC I-IV | S + C | 30.0% | 188 |
| Weber 2001 [14] France 2001 | microscope | quantitative | mouse monoclonal | formalin | 50.2% 61(29–80) | CRC I-IV | S + C | 50.0% | 221 |
| Xi [15] China 2011 | microscope | semi-quantitative | mouse monoclonal | formalin | 59.7% 62(20–81) | CRC I-IV | S | score > 5 | 201 |
| Ivanec [16] Slovenia 2014 | microscope | semi-quantitative | mouse monoclonal | formalin | 28.0% 62(27–78) | CRC I-IV | S + C | 50.0% | 98 |
| Rosati [17] Italy 2004 | microscope | semi-quantitative | mouse monoclonal | formalin | 55.3% 66(29–79) | CRC I-IV | S + C | 10.0% | 103 |
| Nishihara [18] Japan 2009 | microscope | quantitative | mouse monoclonal | formalin | 51.2% 68(33–94) | CRC I-IV | S | 40.0% | 201 |
| Scopa [19] Greece 2003 | microscope | semi-quantitative | mouse monoclonal | formalin | 64.0% 66(25–82) | CRC I-IV | S + C | 5.0% | 117 |
| Visca [20] Italy 1999 | microscope | semi-quantitative | mouse monoclonal | formalin | 90% 64 | CRC I-IV | S | 20.0% | 100 |
| Handa [21] Japan 1999 | microscope | semi-quantitative | mouse monoclonal | formalin | 65.8% 66(39–90) | CRC I-IV | S | 19.5% | 73 |
| Bahnassy [22] Egypt 2004 | microscope | quantitative | NR | formalin | 50.3% 61(56–70) | CRC I-IV | S | 11.5% | 60 |
| Buglioni [23] 1999 | microscope | quantitative | NR | formalin | 90.0% 64(26–78) | CRC I-IV | S | 25.0% | 171 |
| Bertolini [24] Italy 2007 | microscope | quantitative | NR | formalin | 31.0% 63.1 ± 10.7 | CRC I-IV | S + C + R | 20.0% | 91 |
| Kimura [25] Japan 2000 | microscope | quantitative | monoclonal | formalin | 69.1% 63.9(45–96) | CRC II-IV | S | 50.6% | 71 |
| Fernandez-Cebrian [26] Spain 2007 | microscope | semi-quantitative | mouse monoclonal | formalin | 78.7% 62(22–83) | CRC I-IV | S | score > 5 | 192 |
| Wu 2012 [27] China 2012 | microscope | semi-quantitative | mouse monoclonal | formalin | 45.0% 63.2 ± 13.9 | CRC II-IV | S + C | 47.9% | 111 |
| Furudo 2001 [28] Japan 2001 | microscope | quantitative | mouse monoclonal | formalin | 18.3% 64 | CRC I-IV | S | 50.0% | 60 |
| Lin [29] China 2008 | microscope | semi-quantitative | mouse monoclonal | formalin | 63.5% 65(35–83) | CRC I-IV | S + C | 30.0% | 54 |
Table 1 Characteristics of studies included in the present meta analysis (Continued)

| Author/Country/Year | Detection method | Quantitative method | Antibody type | Fixed method | Positive\(^a\) Rate(%) | Median age (Range) | History Type/Stage | Therapy Value | Cut off Value | Patient Number |
|---------------------|------------------|---------------------|---------------|--------------|------------------------|-------------------|-------------------|---------------|---------------|----------------|
| Ajani [31] America 2010 | microscope | quantitative | monoclonal | formalin | 50.0% | NR | RC | I-IV | C + R | median | 30 |
| Li [32] China 2009 | microscope | quantitative | monoclonal | formalin | 78.8% | 65(22–95) | CC | I-IV | S | 10.0% | 203 |
| Hu 2009 [33] China 2009 | microscope | quantitative | NR | formalin | 30.7% | 56.6(34–75) | CRC | NR | S | 10.0% | 52 |
| Roxburgh [34] Scotland 2013 | Tissue microarrays | quantitative | mouse monoclonal | NR | 61% | NR | CRC | I-III | S + C | 15% | 230 |
| Shin 2014 [35] Korea 2014 | microscope | quantitative | NR | NR | 74.4% | 63(30–87) | CRC | I-IV | S + C | 50.0% | 266 |
| Nash [36] America 2010 | microscope | quantitative | NR | formalin | 66.0% | 63(19–84) | CRC | I-IV | S + C | 50.0% | 188 |
| Prall [37] Germany 2004 | Tissue microarrays | quantitative | mouse monoclonal | NR | 50.0% | 64.8(29–90) | CRC | I-IV | S + C + R | median | 149 |
| Zlobec [38] Switzerland 2008 | Tissue microarrays | quantitative | NR | NR | 42.5% | 68.7(36–96) | RC | I-IV | S + C + R | 15.0% | 482 |

\(^a\)positive status was defined according to cut off value

NR not reported, OS overall survival, DFS disease free survival, S surgery, C chemotherapy, R radiotherapy, CRC colorectal cancer, CRCLM colorectal liver metastases, CC colon cancer, RC rectal cancer, NOS Newcastle-Ottawa Quality Assessment Scale, IHC immunohistochemistry

Fig. 2 Forest plot diagrams of Ki-67 expression for overall survival in univariate analysis
Contour-enhanced funnel plot

The contour-enhanced funnel plot can help to ascertain whether funnel plot asymmetry due to publication bias [40]. Contour lines which indicated conventional milestones in levels of statistical significance (0.05, and 0.1) are added to funnel plots. If the dummy studies lie in areas of statistical nonsignificance, then this provides evidence of the possibility that the funnel plot asymmetry is due to publication bias. Conversely, if the dummy studies are in areas of high statistical significance ($P > 0.05$), this will support that the cause of the funnel plot asymmetry may be caused by factors other...
than publication bias, such as poor methodological quality or true heterogeneity.

To explore the potential causes of funnel plot asymmetry for OS in univariate analysis, a contour-enhanced funnel plot by using the trim-and-fill method was performed. The dummy studies were indicated by red triangles and the genuine studies were indicated by green dots (Fig. 6). This trim-and-fill method added 9 dummy studies to balance the funnel plot and all 9 dummy studies were in areas of high statistical significance, suggesting that the publication bias might not the major cause of funnel plot asymmetry in the present meta analysis.

Meta-regression analysis and subgroup analysis

Subgroup analysis (Table 2) suggested that cut off value, treatment and pathologic type might contribute to the clinical heterogeneity. The pooled HRs were 0.743 (95% CI 0.332–1.662, $I^2 = 20.7\%$, $P = 0.261$) for rectal cancer (RC) patients and 1.902(0.695–5.198, $I^2 = 92\%$, $P = 0.001$) for colorectal liver metastases (CRCLM) patients. Further meta-regression analysis did not find any significant source of heterogeneity (all $P > 0.05$).

According to the results of subgroup analysis, we further explored the association between Ki-67 expression and overall survival with different cut off values (Fig. 7). The pooled HR (1.97, 95% CI 1.37–2.83) in group with cut off value (20–30%) was significantly higher than that in other groups, indicating that cut off value (20–30%) might be more suitable for clinical practice than other cut off values.

Sensitivity analyses

All studies were sequentially removed to explore that whether any individual study had a significant influence to the pooled HRs. The pooled HRs of sensitivity meta analysis ranged from 1.49(95%CI: 1.13–1.96) to 1.65 (95%CI: 1.26–2.15) for OS in univariate analysis, demonstrating that the pooled HRs were not significantly affected by any individual study (Table 3).

Cumulative meta analysis of Ki-67 expression

We further performed cumulative meta analysis to assess the stability of Ki-67 expression for OS in univariate analysis (Fig. 8a), OS in multivariate analysis (Fig. 8b) and DFS in univariate analysis (Fig. 8c). The pooled HRs of cumulative meta analysis ranged from 1.48(95%CI 1.1–2.0) to 2.14 (95%CI 1.64–2.79) for OS in univariate analysis since 1999, demonstrating that performance of Ki-67 expression for OS in CRC patients was stable and reliable.

Discussion

The present meta analysis showed that high Ki-67 expression is significantly correlated with poor prognosis in CRC patients. The pooled hazard ratios were 1.54(95% CI 1.17–2.02, $P = 0.005$) for overall survival and 1.43(1.12–1.83, $P = 0.008$) for disease free survival in univariate analysis. After adjustment of other prognostic factors, the pooled HR was 1.50(95% CI 1.02–2.22, $P = 0.03$) for overall survival in multivariate analysis.

Some previous studies have reported that high Ki-67 expression significantly predicted poor OS in CRC patients [26–30]. The conclusion of the present meta
analysis was consistent with that of these previous studies. Although the mechanism of Ki-67 expression on tumor prognosis was still uncertain, some clinical studies have provided interesting evidences for utility of Ki-67 expression as a predictive tool for clinical prognosis in patients with different tumors, for example prostate cancer, thyroid carcinoma and renal clear cell [41–43]. Interestingly, there were three studies indicated that there was a negative association between the high Ki-67 expression and the overall survival of CRC patients. Salminen et al. [8] reported that HR of high Ki-67 expression was 0.63 (95% CI 0.41–0.98) for overall survival in univariate analysis, whereas it was

![Fig. 6 Contour-enhanced funnel plot for overall survival in univariate analysis](image)

| Term            | Subgroup | Number of study | Pooled HR   | 95% CI    | P     | Heterogeneity | P     |
|-----------------|----------|----------------|-------------|-----------|-------|---------------|-------|
| Location        | Asian    | 10             | 1.89        | 1.23–2.90 | 0.003 | 70.9          | 0.001 |
|                 | Non-Asian| 11             | 1.282       | 0.936–1.758| 0.122 | 77.4          | 0.001 |
| Cut off value   | ≥40%     | 9              | 1.5         | 1.01–2.24 | 0.044 | 84.9          | 0.001 |
|                 | <40%     | 10             | 1.71        | 1.33–2.21 | 0.001 | 0             | 0.529 |
| Patient number  | ≥100     | 13             | 1.381       | 0.994–1.918| 0.054 | 75.7          | 0.001 |
|                 | <100     | 8              | 2.008       | 1.131–3.564| 0.017 | 79.4          | 0.001 |
| Positive rate   | ≥50%     | 14             | 1.513       | 1.119–2.046| 0.007 | 64.2          | 0.001 |
|                 | <50%     | 6              | 1.589       | 0.911–2.773| 0.103 | 86.5          | 0.001 |
| Pathologic type | CRC      | 14             | 1.526       | 1.135–2.05 | 0.005 | 62.8          | 0.001 |
|                 | CC       | 2              | 2.257       | 1.418–3.594| 0.001 | 77.7          | 0.001 |
|                 | RC       | 2              | 0.743       | 0.332–1.662| 0.469 | 20.7          | 0.261 |
|                 | CRCLM    | 3              | 1.902       | 0.696–5.198| 0.21  | 92.6          | 0.001 |
| Treatment       | S        | 12             | 1.518       | 1.038–2.221| 0.031 | 74.6          | 0.001 |
|                 | S + C    | 7              | 1.477       | 0.933–2.338| 0.096 | 82.3          | 0.001 |
|                 | S + C + R| 2              | 2.216       | 1.227–4.002| 0.008 | 0             | 0.945 |
| Extraction method | Curve   | 7              | 1.261       | 0.915–1.736| 0.001 | 57.3          | 0.029 |
|                 | Reported | 14             | 2.095       | 1.491–2.02 | 0.156 | 72.5          | 0.001 |

HR hazard ratio, CI confidence interval, CRC colorectal cancer, CRCLM colorectal liver metastases, CC colon cancer, RC rectal cancer, S surgery, C chemotherapy, R radiotherapy

*positive status was defined according to cut off value
1.26 (95% CI 0.67–2.39) for overall survival in multivariate analysis. Xi et al. [15] reported that HR of high Ki-67 expression was 0.34 (95% CI 0.16–0.73) for overall survival by using a special scoring system of Ki-67 expression: The intensity of positivity was scored as follows: 0, negative; 1, weak; 2, moderate; 3, strong. The extent of positivity was scored according to the percentage of cells showing positive staining: 0, < 5%; 1, 5–25%; 2, 25–50%; 3, 50–75%; 4, > 75%. The final score was determined by multiplying the intensity of positivity and the extent of positivity scores and final score > 5 was defined as positive. Ivanecz et al. [16] selected 98 patients for prognostic study from 406 patients by using the following criteria: first liver resection for CLM; potentially curative R0 resection; follow-up period of surviving patients > 5 years. The special inclusion criteria might influence the baseline characteristics and weaken the credibility of study conclusion.

There was significant heterogeneity in the present meta analysis. The current meta analysis might exist several potential sources of heterogeneity as follows: First, the heterogeneity caused by different cut off values of Ki-67 expression was inevitable. The subgroup analysis have revealed the influence of different cut off values on heterogeneity. Second, the subgroup analysis suggested that different treatment has a significant effect on heterogeneity and might be a potential source of heterogeneity. Third, obvious difference of the pooled HRs between rectal cancer subgroup and colorectal liver metastases subgroup suggested that further study is need to clarify the exact influence of different pathologic type on clinical prognosis.

Publication bias is an important influence factor for interpreting the conclusion. The funnel plot for OS in univariate analysis was obviously asymmetrical and Egger’s test also demonstrated significant publication bias ($P = 0.008$). The funnel plot asymmetry might be caused by different reasons as follows: Firstly, publication bias was one important source of funnel plot asymmetry, for example location bias, citation bias, language bias, time lag bias, multiple publication bias and outcome reporting bias. Secondly, the true heterogeneity in intervention effects might be one resource of funnel plot asymmetry. The true heterogeneity might be caused by differences in the intensity of interventions or differences in underlying risk between studies with different...
sizes. Thirdly, poor methodological quality was another source of funnel plot asymmetry. Smaller studies were often performed with poorer methodological quality than larger studies. Therefore, smaller studies with poorer methodological quality also tended to display larger influences than larger studies. Finally, funnel plot asymmetrical might be caused by chance. We conducted a contour-enhanced funnel plot using the trim-and-fill method to explore the potential cause of funnel plot asymmetry. Given that all 9 dummy studies were in areas of high statistical significance, the publication bias might not the major cause of funnel plot asymmetry in the present meta analysis.

Although significant heterogeneity existed in the present meta analysis, sensitivity analysis demonstrated that the relation between high Ki-67 expression and poor prognosis of CRC patients did not change after removing any individual study. Meanwhile, cumulative meta analysis also showed that performance of high Ki-67 expression for poor prognosis in CRC patients was stable and reliable.

The present meta analysis had several strengths: Firstly, there were 34 eligible studies and 6180 CRC patients in the present study, which could significantly increase persuasiveness of the conclusion. Secondly, Ki-67 expression in 34 eligible studies was all measured with IHC method, leading to less clinical heterogeneity. Thirdly, studies published in Chinese were also included in this meta analysis as English literature, increasing the representation of the study population.

The conclusion of the present meta analysis should be interpreted cautiously for reasons as follows: Firstly, most studies defined positive status of Ki-67 expression according to different cut off values, which might result in clinical heterogeneity. Hence, it is necessary to determine an universal cut off value for future studies. Secondly, heterogeneity was inevitable due to different baseline characteristics, such as tumor stage and treatment, which might increase clinical heterogeneity and reduce the persuasiveness of the conclusion. Thirdly, some survival data were extracted from survival curves indirectly. Although the method for extracting HR and 95% CI from survival curves is widely accepted, we could not completely eliminate the sources of information inaccuracy in the process of information extraction. Fourthly, Ki-67 expression in the selected studies was measured by IHC method. This method might suffer from important variation among different studies and was not enough precise for a rigorous and precise conclusion. Fifthly, several genes

| Study omitted | HR    | Lower value of 95% CI | Upper value of 95% CI |
|--------------|-------|-----------------------|----------------------|
| 1            | 1.6470636 | 1.2618686           | 2.1498425            |
| 2            | 1.6286341 | 1.2338084           | 2.1496494            |
| 3            | 1.6086709 | 1.2262427           | 2.110367             |
| 4            | 1.5804045 | 1.1893544           | 2.1000286            |
| 5            | 1.5853532 | 1.1829669           | 2.1246113            |
| 6            | 1.5744233 | 1.1822454           | 2.0966956            |
| 7            | 1.5586835 | 1.1753646           | 2.0670133            |
| 8            | 1.5452128 | 1.1628222           | 2.0533514            |
| 9            | 1.5309644 | 1.158212            | 2.0236814            |
| 10           | 1.5309644 | 1.158212            | 2.0236814            |
| 11           | 1.5234994 | 1.150002            | 2.0183012            |
| 12           | 1.5333095 | 1.1640916           | 2.0196332            |
| 13           | 1.5163125 | 1.145869            | 2.0065153            |
| 14           | 1.5121837 | 1.1424872           | 2.00151              |
| 15           | 1.5081896 | 1.140373            | 1.9947045            |
| 16           | 1.5015592 | 1.136433            | 1.9839972            |
| 17           | 1.5038622 | 1.1387317           | 1.9860705            |
| 18           | 1.4894764 | 1.1326977           | 1.9586338            |
| 19           | 1.4916791 | 1.1332084           | 1.9635456            |
| 20           | 1.4770394 | 1.1257362           | 1.9379722            |
| 21           | 1.4908768 | 1.1371239           | 1.9546804            |
| Combined     | 1.5392004 | 1.1725948           | 2.0204233            |

HR hazard ratio, CI confidence interval
Fig. 8 Cumulative meta analysis of Ki-67 expression. a Overall survival in univariate analysis; b Overall survival in multivariate analysis; c Disease free survival in univariate analysis.
have been reported to be related with the prognosis of CRC patients including KRAS, Cyclin-dependent kinase 2 (cdc2), and Cathepsin D [5, 35, 36]. Interestingly enough, Nash reported that patients with KRAS mutant tumors that expressed high levels of Ki-67 had significantly worse DFS than patients with KRAS wild-type tumors that expressed low level of Ki-67, indicating that combined application of these predictive genes might provide more valuable prognostic information and further improve the clinical prediction of prognosis for CRC patients.

In conclusion, the present meta analysis demonstrated that high Ki-67 expression significantly predicts poor overall survival and disease free survival. High Ki-67 expression may serve as a valuable predictive biomarker for poor prognosis in CRC patients.

Abbreviations
C: Chemotherapy; CC: Colon cancer; CI: Confidence interval; CRC: Colorectal cancer; CRCLM: Colorectal liver metastases; DFS: Disease free survival; HR: Hazard ratio; IHC: Immunohistochemistry; NOS: Newcastle-Ottawa Quality Assessment Scale; NR: Not reported; OR: Odds ratio; OS: Overall survival; R: Radiotherapy; RC: Rectal cancer; S: Surgery

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Availability of data and materials
The datasets supporting the conclusions of the present study were included in the article.

Authors’ contributions
ZZ, ZM and LZ designed the study. ZZ, YF, HW, LX, and LZ performed the research; ZZ and ZM collected and analyzed the data; ZZ and ZM wrote the paper; LZ amended the article. ZZ act as the submission guarantor and takes responsibility for the integrity of the work as a whole, from inception to published article. All authors reviewed the manuscript. All authors read and approved the final manuscript.

Ethics approval and consent to participate
Not applicable.

Consent for publication
All the authors have agreed to publish article in the current magazine.

Competing interests
The authors declare that they have no competing interests.

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