1. INTRODUCTION

When an extrasolar planet eclipses its host star, the event is referred to as a transit or primary eclipse. During a transit, observers can detect dips in starlight caused by the planet obscuring a portion of the stellar disk when it passes in front of the star. The transit method applies to those systems where the orbital inclination of a planet is close to 90° (i.e., edge-on) with respect to the observer’s sky plane. The first transit observations were made by Charbonneau et al. (2000). As of 2012 August, more than 200 planets (http://exoplanet.eu) have been detected using the transit method. The search for new exoplanetary systems via transits has advanced to space-based missions with the launch of the European Space Agency’s CoRoT satellite (Fridlund et al. 2006) and NASA’s Kepler mission (Borucki et al. 2010). The objective of transit detection and follow-up is mainly to catalog and improve measurements of system parameters. Studying the characteristics of extrasolar planetary systems is crucial for developing theories of planet formation that can adequately explain the origin and evolution of all planetary systems (including our own).

The target discussed in this paper, TrES-3b, is a hot-Jupiter with one of the shortest orbital periods known (P = 1.3 days; O’Donovan et al. 2007) among the exoplanets. The planet orbits a G-type star (Teff = 5720 K), and has a mass and radius of M_p = 1.92 M_Jup and R_p = 1.29 R_Jup, respectively (O’Donovan et al. 2007). Due to its large impact parameter, TrES-3b’s transit has more of a v-shape than the typical u-shape. However, the transit is not grazing and the ability of various observers to consistently measure the transit parameters seems to indicate that the planet’s disk completely enters the stellar disk during the transit (O’Donovan et al. 2007; Sozzetti et al. 2009; Gibson et al. 2009). A blended eclipsing binary, which could account for the v-shape, has been ruled out from the radial velocity analysis (O’Donovan et al. 2007). However, the v-shape of the transit makes measurements of transit properties challenging at shorter wavelengths, where stellar limb-darkening further degrades the trapezoidal u-shape of the transit. The level of insolation received on the dayside of TrES-3b due to its proximity to its host star should place it in the class of warm hot-Jupiters with a temperature inversion in its atmosphere (Hubeny et al. 2003; Burrows et al. 2007; Fortney et al. 2008). However, secondary eclipse observations from the ground and space have shown that the evidence for an inversion layer is not strong (de Mooij & Snellen 2009; Fressin et al. 2010). TrES-3b’s lack of an inversion layer is consistent with the idea that UV radiation from chromospherically active stars (like TrES-3b’s parent star) systematically destroys absorbers in the upper planetary atmosphere, thereby preventing the formation of an inversion layer (Knutson et al. 2010).

The ultra close-in orbit of TrES-3b also makes it a great target for follow-up transit monitoring. It has been suggested that planets with very short periods are likely to be falling into their host stars (Jackson et al. 2009). Some have looked for transit timing variations (TTVs) indicative of such orbital decay (e.g., OGLE-TR-56b; Adams et al. 2011). TrES-3b is similar to OGLE-TR-56b in many regards, including the orbital period (<1.5 days), size (∼1 R_Jup), and impact parameter (b > 0.8). Though there has been no strong evidence for TTVs for OGLE-TR-56b, it is one object that warrants long-term study (Adams et al. 2011). Transit monitoring of TrES-3b by other teams has so far confirmed its linear orbital ephemeris (Soffetti et al. 2009; Gibson et al. 2009). A search for additional transiting companions using the EPOXI mission has yielded no detections, since the likelihood of detecting a planetary sibling in a coplanar orbit with TrES-3b is lower for an outer planet given the inclination of the system (Ballard et al. 2011); inner planets are not expected since TrES-3b is already quite close to its parent star.

In this paper we report observations of 11 transits of TrES-3b, taken as part of the Apache Point Observatory Survey of Transit Lightcurves of Exoplanets (APOSTLE). The APOSTLE program is a follow-up transit monitoring program designed to obtain high-precision relative photometry on known transiting
systems in order to refine measurements of system parameters and transit times (see, e.g., GJ 1214b; Kundurthy et al. 2011). In Section 2 we outline our observations, and in Section 3 we describe the data reduction. The two sections that follow, Sections 4 and 5, outline the transit model and the Markov Chain Monte Carlo (MCMC) analyzer, respectively. In Section 6 we present our estimates of the system parameters for TrES-3b and in the Sections 6.1 and 6.2 we present results from our study of transit depth variations (TDVs) and TTVs. Finally, in Section 7 we summarize our findings.

2. OBSERVATIONS

TrES-3 was observed by APOSTLE over a time span of two years between the summer of 2009 and the fall of 2011. All observations of TrES-3 were carried out using AGILE a high-speed frame-transfer photometer (Mukadam et al. 2011), on the ARC3 3.5 m telescope at Apache Point, New Mexico. The AGILE CCD has no dead time, as the charge is transferred to an adjoining array for read-out. All observations for TrES-3 were made using AGILE’s medium-gain, slow read-out (100 kHz) mode, with the charge read-out at 45 s intervals. The filter used for the observations was the r band similar to the SDSS4 r filter (with central wavelength, $\lambda_0 = 626$ nm; Fukugita et al. 1996). This observing filter is bluer than typical filters used for transit observations since AGILE is a blue sensitive CCD that is affected by a strong fringe pattern at longer wavelengths (Mukadam et al. 2011). The summary of the 11 r-band observations is given in Table 1. Observations were made by adjusting the focus on the secondary mirror to smear the stellar point-spread functions (PSFs) across multiple pixels, which minimizes the systematics caused by pixel-to-pixel wandering of the PSF. The long read-out (exposure time) also allowed for a greater count rate that maximized the signal-to-noise per image. The count rate was kept below AGILE’s nonlinearity limit of $\sim 52$ k ADU and well below its saturation level of 61k ADU by small adjustments to the telescope’s secondary focus.

The parent star of TrES-3b, GSC 03089-00929 (TrES-3), is a G-type star with a Johnson R magnitude of 12.2. The comparison star used for the relative photometry was USNO-B1.0 1275-0332540 situated $\sim 90^\circ$ away, with Johnson R of 12.9 (Monet et al. 2003), which was the next brightest star in AGILE’s field of view (FOV). Several studies that use detectors with larger FOVs than AGILE typically use many ($N \gg 1$) comparison stars. For relative aperture photometry, lightcurve precision is limited by the signal to noise achieved from aperture extraction on the faintest star in the set. The small FOV of AGILE meant that most other comparison stars in the field were too faint to provide the adequate signal to noise on the final lightcurve. PSF photometry is a solution to the problem of using faint comparison stars since the stellar and background components can be constrained accurately. However, due to the difficulty in modeling the complex defocused PSF of APOSTLE observations, we did not opt for this route. We used a circular aperture for photometry (see details in Section 3) and hence the target and brightest companion produced the best results. Our uncalibrated differential photometry shows that TrES-3 was the brighter of the two by a factor of $\sim 1.8$ in the $r'$ band. We also note variability in the uncalibrated flux between TrES-3 and its reference star, with the maximum difference being $\sim 20\%$ between the highest (7 UTD 2011 March 24) and lowest (3 UTD 2010 March 22) values. The observations were made over a variety of observing conditions (Column “Obs. Conditions” in Table 1). The listed nights include both complete and partial transits (where data were lost due to poor weather or issues with the instrument). Some of these partial transits include UTD 2009 May 14 (1), 2009 June 13 (2), and 2010 June 2 (10). Small portions of the in-eclipse data were lost for the transits on UTD 2010 March 22 (3) and UTD 2011 April 27 (8), and the night of UTD 2011 May 14 (9) had exceptionally poor observing conditions (seeing $>2^\prime$). Even though these data will affect the fit, we include them in the analysis since they can be used to determine transit times and other system parameters.

3. APOSTLE PIPELINE

The APOSTLE project used a customized data reduction pipeline, written in the Interactive Data Language to process data from AGILE. The pipeline performs standard image processing steps like dark subtraction and flat-fielding, but also implements nonlinearity corrections unique to AGILE. The pipeline also creates an uncertainty map of the processed images by propagating pixel-to-pixel errors through each step of the reduction. In addition to the photon counting errors and read-noise from the raw images, the pipeline propagates the variance on the

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Table 1

| T. No. | UTD Date | Obs. Condition | Filter | Exposure (s) | Phot. Aperture (pixels) | rms (ppm) | % Rej. | Flux Normalization | Error Scaling |
|--------|----------|----------------|--------|--------------|-------------------------|-----------|--------|-------------------|--------------|
| 1      | 2009 May 14 | Clear | r     | 45           | 30                      | 777       | 2%     | 1.9179            | 0.5618       |
| 2      | 2009 Jun 13 | Poor weather | r'    | 45           | 22                      | 610       | 13%    | 1.9083            | 0.5135       |
| 3      | 2010 Mar 22 | Clear | r     | 45           | 20                      | 962       | <1%    | 1.7788            | 0.7436       |
| 4      | 2010 May 16 | Clear | r'    | 45           | 21                      | 877       | 1%     | 1.8954            | 0.7769       |
| 5      | 2010 Jun 02 | Poor weather | r'   | 45           | 30                      | 703       | <1%    | 1.8670            | 0.2287       |
| 6      | 2010 Oct 12 | Clear | r'    | 45           | 24                      | 896       | <1%    | 1.8832            | 0.7804       |
| 7      | 2011 Mar 24 | Clear | r     | 45           | 19                      | 644       | 1%     | 1.9809            | 0.5644       |
| 8      | 2011 Apr 27 | Poor weather | r'   | 45           | 14                      | 1022      | 1%     | 1.8932            | 0.9787       |
| 9      | 2011 May 14 | Poor weather | r'   | 45           | 15                      | 2960      | <1%    | 1.8887            | 2.8965       |
| 10     | 2011 Jun 21 | Clear | r     | 45           | 26                      | 1221      | <1%    | 1.9227            | 0.9683       |
| 11     | 2011 Aug 24 | Clear | r'    | 45           | 19                      | 1240      | 7%     | 1.8016            | 0.7597       |

Notes. (1) Transit Number, (2) Universal Time Date, (3) Observing Conditions, (4) Observing Filter, (5) Exposure (seconds), (6) Optimal Aperture Radius (pixels), (7) Scatter in the residuals, (8) % frames rejected due to saturation or other effects, (9) Flux normalization between the target and comparison star, (10) The factor by which the photometric errors were scaled.
master dark and master flat during the reduction. Errors were also propagated for those pixels where the counts exceeded the nonlinearity threshold (∼52 k counts) using the uncertainties in an empirically derived nonlinearity correction function. Frames where pixels inside a photometric aperture exceeded AGILE’s saturation limit of 61 k were rejected. Images at the other extreme, where the stars were obscured by clouds, and resulted in low signal-to-noise measurements were also rejected (i.e., where photometric errors were >5000 ppm). The fraction of rejected frames per night is listed in Column 9 “%Rej” in Table 1.

We used SExtractor (Bertin & Arnouts 1996) to derive initial centroids of our defocused stars. This software allows the use of a customized PSF kernel, so we used a “donut”-shaped detection kernel for our defocused data. Coordinates obtained from SExtractor were then used for circular aperture photometry with the PHOT task in IRAF’s NOAO.DIGIPHOT.APPHOT package. We derived flux estimates from a range of circular apertures with radii between 5 and 50 pixels, at intervals of 1 pixel, by simply summing the counts in these apertures. An outlier-rejected global median on the frame was used as the sky estimate, which is removed to derive the instrumental flux of the stars. To derive photometric errors, we extracted counts from the error frames using the same centroids and apertures used for photometry on the target frames. The lightcurves are generated by dividing the instrumental flux of the target star by the comparison star and then dividing the entire lightcurve by the median out-of-eclipse flux level. At this stage there may still be systematic trends in the lightcurve that have not been removed by reduction; for example, differential extinction due to airmass variation or photometric variation due to centroids wandering over pixels of varying sensitivities (e.g., due to small imperfections in the flat-fielding). Thus, for each image we also extracted a set of nuisance parameters which are used to compute a correction function (i.e., detrending function). For the TrES-3 data we found that (1) the airmass, (2) the global median sky, (3) the centroid positions of the target star, and the sum of counts in the photometric aperture for (4) the master dark, and (5) the master flat showed trends that corresponded to trends seen in the lightcurves. The airmass was derived from the image headers, while the global sky and the centroid positions are derived from the photometry on the science frames. The sum of counts in the master dark and master flat are derived from photometry on the master dark and flat using the centroids and apertures used for photometry on the science frames. The correction function ($F_{\text{cor}}$) is modeled as a linear sum of nuisance parameters as described by the following equation:

$$F_{\text{cor},i} = \sum_{k=1}^{N_{\text{nus}}} c_k X_{k,i},$$

where $X_{k,i}$ are the nuisance parameters, $c_k$ are the corresponding coefficients. The index $k$ counts over the number of nuisance parameters $N_{\text{nus}}$, and the index $i$ denotes the transit number. The detrending coefficients are chosen by minimizing the $\chi^2$ between the observed data ($O$), a model function ($M$), and correction function,

$$\chi^2 = \sum_j \frac{(O_j - M_j - F_{\text{cor},j})^2}{\sigma_j^2},$$

here $j$ is the index that counts over the total number of data points ($N_{\text{all}}$), when all transits are stacked. During photometry we use a set of trial model parameters (based on values from the literature) to remove the transit lightcurve from the data. We use a linear least-squares minimizer to fit for the coefficients of $F_{\text{cor},i}$ (Equation (1)), which removes any correlated trends that remain in the model-subtracted lightcurve. The resulting residuals are used for selecting the optimal aperture from which to extract the photometry. The aperture where the scatter in the residuals is minimized was chosen as the optimal aperture. Smaller apertures do not completely sample the flux from the stars, and at larger apertures one loses signal to noise due to the fact that one accumulates fewer counts from the star, compared to sky counts. The optimal aperture typically fell between 5 and 50 pixel radii. We present the complete list of optimal apertures in Column 6 in Table 1. The sizes of the typical apertures were always a few times larger than the typical half-width at half-maximum of the stellar PSFs. The photometric precision, which is simply the scatter in the residuals, is also listed in Column 7 in Table 1.

The 11 transits of TrES-3b are shown in Figure 1 in normalized flux ratio units. The horizontal axis shows time from the mid-transit time in days, computed by subtracting the appropriate mid-transit time for each transit from the best-fit values in the Fixed LDC chain.
Table 2

| T. No. | T-T0 | Norm. Fl. Ratio | Err. Norm. Fl. Ratio | Model Data |
|--------|------|-----------------|----------------------|------------|
| 1      | −0.0490181 | 0.999407424 | 0.0000934904 | 1.000000000 |
| 2      | −0.0479764 | 1.000889327 | 0.0000936546 | 1.000000000 |
| 3      | −0.0474555 | 1.001329344 | 0.0000936130 | 1.000000000 |
| 4      | −0.0469547 | 1.000241971 | 0.0000941452 | 1.000000000 |
| 5      | −0.0455226 | 1.000658752 | 0.0000936362 | 1.000000000 |
| 6      | −0.0444809 | 0.998849209 | 0.0000923458 | 1.000000000 |
| 7      | −0.0434392 | 0.998014984 | 0.0000914333 | 1.000000000 |
| 8      | −0.0429184 | 0.996495746 | 0.0000917120 | 1.000000000 |
| 9      | −0.0423975 | 0.99481076 | 0.0000914379 | 1.000000000 |
| 10     | −0.0418767 | 1.000283709 | 0.0000988789 | 1.000000000 |

Notes. (1) Transit Number, (2) Time Stamps - Mid-transit Times (BJD), (3) Normalized Flux Ratio, (4) Error on Normalized Flux Ratio, (5) Model Data.

(This table is available in its entirety in a machine-readable form in the online journal. A portion is shown here for guidance regarding its form and content.)

4. MULTI TRANSIT QUICK

We developed a transit model called Multi Transit Quick (MTQ) in PYTHON, which is based on the analytic lightcurve models presented in Mandel & Agol (2002), and the PYTHON implementation of some of its functions (from EXOFAST by Eastman et al. 2012). The set of transit parameters used by MTQ includes the transit duration ($T_\tau$), the limb-crossing duration ($t_G$), and the times of mid-transit ($T_0$). The parameters $T_\tau$ and $t_G$ are the same as $T$ and $\tau$ from Carter et al. (2008). The shape of the transit is characterized by the transit depth ($D$) and the stellar limb-darkening parameters $v_1$ and $v_2$. The parameter $D$ simply represents the maximum depth of the transit trough at conjunction, which depends on the ratio of the area of the disks of the planet to the star ($R_p^2/R_\star^2$) and the ratio of the disk-averaged stellar intensity at conjunction $I(b)$ to the unobscured disk-averaged intensity at conjunction for an impact parameter of zero, $I(0)$. The transit depth $D$ is

$$D = \frac{R_p^2}{R_\star^2} \frac{I(b)}{I(0)},$$

the term $I(b)/I(0)$ can be replaced by the quadratic limb-darkening profile (Mandel & Agol 2002) to give

$$D = \frac{R_p^2}{R_\star^2} \left(1-u_1(1-\sqrt{1-b^2})-u_2(1-\sqrt{1-b^2})^2\right)/(1-\frac{u_1}{2}-\frac{u_2}{6}),$$

where $u_1$ and $u_2$ are the quadratic limb-darkening coefficients (same as $\gamma_1$ and $\gamma_2$ in Mandel & Agol 2002). The limb-darkening parameters $v_1$ and $v_2$ in MTQ are linear combinations of the quadratic limb-darkening coefficients $u_1$ and $u_2$, with $v_1 = u_1 + u_2$ and $v_2 = u_1 - u_2$. These linear combinations were used while fitting the lightcurves, since it is known that directly fitting for limb-darkening coefficients results in strongly anti-correlated error distributions for various transit parameters (Brown et al. 2001) and severely hinders the chance of Bayesian techniques from converging to accurate values (more in Section 5).

We wrote two versions of MTQ, one designed to fit transits observed with different filters (Multi-Filter), and the other designed to look for variations in the transit depth (Multi-Depth). If one makes transit observations using filters of different wavelength, the resulting set of transit lightcurves must be analyzed using a model that accounts for the different limb-darkening profiles, and differing transit depths extant in the data. The standard set of parameters used for Multi-Filter version of MTQ is $\Theta_{\text{Multi-Filter}} = \{T_\tau, t_G, D_{1...N_T}, v_1, j...N_F, v_2, j...N_F, T_{1...N_T}\}$, where $T_{\tau}$ are the transit times and $v_1$ and $v_2$ are the limb-darkening parameters described in the previous paragraph. The subscripts $i...N_T$ and $j...N_F$ are used to denote multiple transits ($N_T$) and multiple filters ($N_F$), respectively. For APOSTLE’s TrES-3b data we only observed using one filter, and the number of transit was 11.

The second version of MTQ is designed to fit for the depths of each individual transit lightcurve separately. Variations in transit depth can arise for several reasons. Commonly invoked sources of such variations are starspots (cool photospheric regions) and faculae (hot photospheric regions; Font et al. 2007; Lanza et al. 2009; Knutson et al. 2011 and many more). The appearance and disappearance of spots on the stellar surface due to rotation or stellar activity cycles would result in variations in the transit depth. Transits occurring across the spotless stellar surface will be shallower than when active regions exist on the visible face of the star (for cool spots). Conversely, if these active regions are hot spots, the transit depths would change in the opposite manner. Other sources for TDVs include planetary oblateness, spin precession (Carter & Winn 2010), planetary rings (Barnes & Fortney 2004), and satellites (Sartoretti & Schneider 1999; Tousinski & Valio 2011). The wide variety of proposed sources of TDV means there may be several degeneracies to resolve if a TDV is indeed detected. Nonetheless, understanding such phenomena may only be possible by establishing statistically significant measurements showing variable transit depths. The multi-filter capability of MTQ can be modified to fit for the depth of each transit as a unique parameter. In this case the set of parameters used is $\Theta_{\text{Multi-Depth}} = \{T_\tau, t_G, D_{1...N_T}, v_1, j...N_F, v_2, j...N_F, T_{1...N_T}\}$, where $D_{i}$ is now fit for each transit instead of each filter. However, the filter corresponding to each depth is still tracked as the model needs to convolve the limb-darkening profile to correctly reproduce the full lightcurve profile. For TrES-3b observations only a single filter was used. In addition, one transit is chosen as the “reference” transit, using which MTQ internally computes several transit parameters such as $R_p/R_\star$, $a/R_\star$, etc. We picked transit 7 since it had high photometric precision, and was well sampled.

In the following sections we describe results from our attempt to fit for transit parameters using both the Multi-Filter and Multi-Depth models. The following section outlines the MCMC analyzer used in conjunction with MTQ.

5. TRANSIT MCMC

MCMC analyzers are now the standard for modeling data on exoplanets (Ford 2005; Holman et al. 2006; Collier Cameron et al. 2007; Kundurthy et al. 2011; Gazak et al. 2012). We developed an MCMC routine called Transit MCMC (TMCMC), which is designed to work in conjunction with MTQ. However, it can also be used to fit other models and data easily. The core TMCMC routine uses the Metropolis–Hastings (M-H) algorithm, and is based on its implementation for astronomical data, as described by Tegmark et al. (2004) and Ford (2005). The Markov chain is computed by making jumps in parameter space and selecting those jumps which tend toward regions of parameter space at lower $\chi^2$. Jumps from the $j$th step in the
chain are made with the following equation:

\[ \theta_{j+1} = \theta_j + G(0, \sigma^2_{\theta}) \mathbf{f}, \]  

(4)

where \( \theta \) and \( \sigma_{\theta} \) are the vectors of model parameters and their associated step-sizes, respectively. The term \( G(0, \sigma^2_{\theta}) \), referred to as the proposal distribution, is a random number drawn from a normal distribution with a mean of 0 and a variance of \( \sigma^2_{\theta} \). It is customary to use a Gaussian proposal distribution but not mandatory. We modify the M-H algorithm slightly by allowing adaptive jump-size adjustments that optimize the sampling rate of the Markov chain integrator. The desired acceptance rates are achieved by adjusting the step-size modifier \( f \) every 100 accepted steps according to

\[ f_{\text{new}} = W(f_{\text{old}}/N_{\text{trials}}), \]

where \( N_{\text{trials}} \) are the number of steps attempted for the last 100 accepted steps and \( W \) is a scaling factor which is 225 or 434 for single-parameter or multi-parameter chains, respectively (as noted in Collier Cameron et al. 2007). Given sufficient time, a converged chain will have traversed parameter space such that the ensemble of the points accurately represent the uncertainty distributions of the parameters in \( \theta \).

5.1. Markov Chains for MTQ

For APOSTLE data sets, we explored system parameters using three different kinds of chains. Two of these were based on the Multi-Filter parameter set \( \theta_{\text{Multi-Filter}} \) described in Section 4. First, with Fixed Limb-Darkening Coefficients, and second, with Open Limb-Darkening Coefficients. The term limb-darkening coefficient will henceforth be abbreviated as LDC. For the Fixed LDC chains (FLDC), the coefficients were simply fixed to values tabulated for the appropriate observing filter (Claret & Bloemen 2011). For the Open LDC chains (OLDC), the limb-darkening parameters \( v_1 \) and \( v_2 \) are allowed to float. It has been shown that measured limb-darkening coefficients from high-precision studies are in good agreement with tabulated values (Brown et al. 2001; Tingley et al. 2006). Another study using spectrophotometry from Hubble Space Telescope STIS showed that any variations in the estimate of transit parameters due to inaccurate limb-darkening are lower than the 1σ uncertainty, with greater disagreement seen at shorter wavelengths (Knutson et al. 2007). Ground-based observations typically cannot achieve the level of photometric precision of space-based studies, and hence constraining the limb-darkening from transit observations is difficult. Subtle inaccuracies in the fit limb-darkening profile can lead to incompatible estimates of system parameters, especially at short wavelengths (Kundurthy et al. 2011). In effect, the purpose of the Fixed LDC and Open LDC chains is simply to compare the differences in the fit limb-darkening coefficients to the tabulated values in the literature. The third type of Markov chain was run on the Multi-Depth parameter set \( \theta_{\text{Multi-Depth}} \) described in Section 4. APOSTLE lightcurves were gathered over a long time baseline, and statistically significant depth variations seen in the data may help shed light on the various phenomena responsible for depth variations (see Section 4).

We applied bounds to several transit parameters and combinations of transit parameters in MTQ (as shown in Table 3). Most of these bounds were simply to check if the parameters had values that were physically realistic. For most parameters it was quite rare that the MCMC chains got close to the bounding limits, since the chains spend most of their time near low \( \chi^2 \) regions; the best-fit parameter values for the TrES-3 system are far from any physical bounds. When fitting for the limb-darkening, however, the degeneracies proved to be very significant, and the Markov chains often strayed to unrealistic values. For example, \( u_1 \) and \( u_2 \) took on values that would suggest limb-brightening rather than limb-darkening, either as a result of noisy data or due to gaps in the ingress or egress portions of the lightcurve. Limb-brightening is considered unphysical for broadband observations. Thus, after every jump (Equation (4)), the vector of proposal parameters \( \theta_{j+1} \) is run through a series of sub-routines in MTQ, to check if any of the conditions in Table 3 are violated. If any one of these conditions are violated, the vector is discarded and a new vector is generated. TMCMC carries out this check until agreeable values emerge, and then continues with the rest of the algorithm.

5.1.1. Executing and Analyzing Chains

By varying the entire vector of model parameters and applying a single step-size modifier (Equation (4)), we run the risk of using mismatched step-sizes and hence undersampling the posterior distributions of some parameters. As mentioned before, well-constructed chains will properly sample posterior distributions given the correct acceptance rate (Gelman et al. 2003). The key to properly constructing a chain is to choose the relative starting step-sizes for the parameter ensemble, such that they all roam high and low probability regions of parameter space at roughly the same rate. Determining a reasonable set of starting step-sizes is done by running short exploratory Markov chains (40,000 steps) for each model parameter (holding all others fixed). If these exploratory chains have not stabilized to the optimal acceptance rate of ~44% (as noted by Gelman et al. 2003 for single-parameter chains) at the end of 40,000 steps, the chain is run until this rate is achieved. The jump sizes near the end of stabilized exploratory chains are then used as the starting steps for the multi-parameter chains. These multi-parameter chains will also be referred to as “long chains” from now on.

For each transiting system, we ran long chains of \( 2 \times 10^6 \) steps from two different starting locations for each model scenario: Fixed LDC, Open LDC, and Multi-Depth/Fixed LDC. After completion we (1) cropped the initial stages of these chains to remove the burn-in phase, where the chain is far from the best-fit region, and (2) we exclude the stage where the chain is far from the optimal acceptance rate of 23% ± 5%, as noted for multi-parameter chains (Gelman et al. 2003). We then run three types of post-processing on the chains after cropping: (1) we compute the ranked and unranked correlations in the chains of every fit parameter with respect to the others. These statistics provide an estimate of the level of degeneracy between parameters in a given model. The next post-processing steps are two commonly used diagnostics to check for chain convergence.

### Table 3

| Bounds | Notes |
|--------|-------|
| \( t_r > 0 \) | Non-zero transit duration |
| \( t_c > 0 \) | Non-zero limb-crossing duration |
| \( D > 0 \) | Non-zero transit depth |
| \((1-b^2) > 0\) | Impact parameters less than 1 (primary condition for transit) |
| \(b/(a/R_\star) \leq 1\) | Ensures real values for orbital inclination |
| \(0 < u_1 < 1\) | Reasonable limb-darkening coefficients* |
| \(0 < u_2 < 1\) | Reasonable limb-darkening coefficients* |

* Applied when parameters were fit (OLDC chains).
namely (2) computing the auto-correlation lengths and (3) the Gelman–Rubin $R^\text{static}$ values (Gelman & Rubin 1992).

In a Markov chain, since a given step is only dependent on the preceding step, sections in the chain may show trends (i.e., are correlated). The correlation length signifies the interval at which sampled points in a Markov chain will be uncorrelated. The effective length is the total number of points in the chain divided by the correlation length. Short correlation lengths (i.e., large effective lengths) indicate that the MCMC ensemble represents a statistically significant and hence more precise sampling of the posterior distribution (Tegmark et al. 2004). Effective lengths >1000 steps are commonly considered to be satisfactory. The $R^\text{static}$ statistic is computed using multiple chains of the same parameter set that have different initial conditions. An $R^\text{static}$ close to 1 (to within 10%) indicates that all chains have converged, cover approximately the same region of parameter space, and that the relevant parameter space has been sufficiently explored. Results from a chain are deemed useful if the auto-correlation and Gelman–Rubin conditions have been met.

### 5.2. Comparing with the Transit Analysis Package

It has been noted by Carter & Winn (2009) that transit lightcurves lacking any significant “defects” or artificial trends can have correlated noise buried within the overall scatter, which is invisible to visual examination. They test a wavelet-based red-noise model on simulated transit lightcurves with and without artificial red-noise and find that models that do not fit for red-noise are subject to inaccuracies in transit parameters on the order of $2\sigma - 3\sigma$ and tend to have underestimated errors by up to 30%. For transit timing studies, poor estimates such as these are cause for concern, since smaller errors and large deviations from the expected time can easily lead to false claims of TTVs. The detrending routine within TMCMC removes long-term trends related to instrumental or other nuisance parameters. We visually examined the residuals of APOSTLE lightcurves (those observed during good conditions), and noted that the data may still have low-level correlated noise, even after detrending. The Transit Analysis Package (TAP; Gazak et al. 2012) implements the red-noise model of Carter & Winn (2009). We ran separate fits of transit parameters using TAP on the detrended lightcurves from our TMCMC fit.

The typical TAP parameter set is $\theta_{\text{TAP}} = \{a/R_\star, t_{\text{orb}}, (R_p/R_\star)_\text{F}, T_{\text{in}}, \sigma_{\text{white}}(\text{white}), \sigma_{\text{red}}(\text{red})\}$, where $\sigma_{\text{white}}(\text{white})$ and $\sigma_{\text{red}}(\text{red})$ are the white-noise and red-noise levels for $N_T$ transits, respectively. The TAP package does not fit for the period using the transit times, and often yields poor estimates of the period, so we fixed the period to that from fitting for the period using the transit times, and often yields poor estimates of the period, so we fixed the period to that from TAP. The period was fixed at 0.03635 days.

One must also note a pitfall of the Carter & Winn (2009) red-noise model. Their fitting function expects the white-noise and red-noise components of a given lightcurve to be stationary, i.e., there are no temporal variations in the Gaussian white-noise level and the red-noise’s power-spectrum amplitude with time. They note that more elaborate noise models may be required to account for the fact that real data do not conform to these requirements. Typically the scatter is increased (due to lower signal to noise) at the portion of the lightcurve where the sky brightness is higher (like data taken close to twilight). Variable observing conditions may alter not just the white-noise properties of a lightcurve but also result in red-noise that is more complex than what can be described by Carter & Winn’s (2009) wavelet model. In addition, at the 1000 ppm level, stellar variability is not understood; it is reasonable to assume that the target and comparison may each contribute to the correlated noise in the lightcurve. In spite of these caveats, the TAP red-noise analysis serves as a good secondary check to results derived using TMCMC (which does not account for red-noise).

### 6. SYSTEM PARAMETERS

As described in Section 5.1.1, we ran three chains on lightcurves of TrES-3b: using the $\theta_{\text{Multi-Filter}}$ parameter set with (1) Fixed LDC, (2) Open LDC, and (3) using the $\theta_{\text{Multi-Depth}}$ parameter set with Fixed LDC. The parameter sets are described in Section 4. All TrES-3b chains use $t_G, t_R$ and include the transit times $T_i$ for the 11 transits. Since all TrES-3b observations were taken in the $r'$ band, the Fixed LDC and Multi-Depth Fixed LDC chains fit for two LDCs for the $r'$ band. The Multi-Depth chain was set up similarly, but with 11 free parameters for the transit depths. Post-processing statistics and other data for these chains are listed in Table 4. The column “$N_{\text{free}}$,” “Chain Length,” “Corr. Length,” and “Eff. Length” list the number of free parameters, the length of the cropped chain, the correlation, and effective lengths, respectively. All chains were run for approximately 2 million steps, but about 100,000 of the initial steps were removed to account for “burn-in” and selection rate stabilization. The Open LDC and Multi-Depth chains both have quite long effective lengths indicating poor Markov chain statistics. Only the Fixed LDC, $\theta_{\text{Multi-Filter}}$ model satisfies the condition of a well-sampled posterior distribution (effective length is >1000).

The final two columns list the goodness of fit (i.e., lowest $\chi^2$ in the MCMC ensemble) and Degrees-of-freedom (DOF) from the respective chain. Parameters from all chains had Gelman–Rubin $R$-statistics close to 1 indicating that the parameter space was covered evenly (though the OLDC and Multi-Depth FLDC chains were not sampled finely enough, based on the auto-correlation data). As noted previously in Kundurthy et al. (2011), the $r'$-band limb-darkening is difficult to characterize using transit models. This result is seen again given the fact that the OLDC chain has the worst auto-correlation statistics. In addition, the fact that TrES-3b has a near-grazing transit makes constraining the transit shape more challenging. The Multi-Depth models also did not converge, which could be due to the fact that several of the transits are not completely sampled during the eclipse event (see Figure 1).



| Chain | Model Vector | $N_{\text{free}}$ | Chain Length | Corr. Length | Eff. Length | $\chi^2$ | DOF |
|-------|--------------|------------------|--------------|--------------|-------------|---------|-----|
| FLDC  | $\theta_{\text{Multi-Filter}}$ | 14 | 1,900,001 | 318 | 5974 | 2503.67 | 2575 |
| OLDC  | $\theta_{\text{Multi-Filter}}$ | 16 | 1,900,001 | 4996 | 380 | 2519.14 | 2573 |
| MDFLDC | $\theta_{\text{Multi-Depth}}$ | 24 | 1,900,001 | 2866 | 662 | 2692.94 | 2565 |
The resulting best-fit parameter estimates are listed in Table 5 for the Multi-Filter models, and in Table 6 for the Multi-Depth models. These tables also list the derived system parameters. The transformation between the MTQ parameters to the derived system parameters is described in Carter et al. (2008) and Kundurthy et al. (2011). Contour plots showing the joint probability distributions (JPDs) for the fit and derived parameters are shown in Figures 2 and 3, respectively.

It is interesting to note that there is a clear degeneracy in the MTQ parameters as seen by the correlations in the posterior distributions of three parameters \((D, t_G, \text{and } t_T)\) shown in Figure 2. From the impact parameter \((b = 0.84)\) we note, as previous studies have, that TrES-3b has a near-grazing transit. The correlation in the JPDs seen in Figure 2 can be understood from Equation (3b) and the following equation (Carter et al. 2008):

\[
b = \sqrt{1 - (t_T/t_G)(R_p/R_*)} .
\]

Rearranging these two expressions gives an expression where \((t_Gt_T) \propto (D/1-b^2)\). Carter et al. (2008) have shown that when \(b \rightarrow 1\) (i.e., the transit is close to grazing), the fraction \(t_G/t_T\) rises and the covariances between \(D, t_G, \text{and } t_T\) rapidly deviate from zero. This degeneracy may be the primary reason why there is disagreement between the estimates of the system parameters for TrES-3b reported by various groups (see Table 7). It may also explain the low statistical significance of the Open LDC and Multi-Depth/Fixed LDC chains, since Markov chains take

\[
\begin{array}{c|c|c}
\text{Parameter} & \text{FLDC} & \text{OLDC} \\
\hline
r^*(\text{days}) & 0.0210 \pm 0.0004 & 0.0212^{+0.0007}_{-0.0005} \\
\tau_r & 0.0383 \pm 0.0002 & 0.0385 \pm 0.0010 \\
D(\nu) & 0.0251 \pm 0.0002 & 0.0255 \pm 0.0003 \\
v_1(\nu) & 0.6767 & 0.7371 \pm 0.1279 \\
v_2(\nu) & 0.3008 & -0.1782 \pm 0.4955 \\
\hline
\end{array}
\]

\[
\begin{array}{c|c|c}
\text{Parameter} & \text{FLDC} & \text{OLDC} \\
\hline
\rho^* & 0.836 \pm 0.003 & 0.837 \pm 0.0008 \\
a/R_* & 5.97 \pm 0.03 & 5.91^{+0.04}_{-0.05} \\
\nu/R_* & 28.71 \pm 0.14 & 28.4^{+0.19}_{-0.26} \text{ day s}^{-1} \\
\rho_* & 2.36 \pm 0.03 & 2.29^{+0.05}_{-0.06} \text{ g/cc} \\
P & (1.3062 \text{ days} +) & -1141 \pm 21 & -1109 \pm 22 \text{ ms} \\
\hline
\end{array}
\]

\[
\begin{array}{c|c|c|c|c}
\hline
\text{Transit Depths} & \text{Value} & \text{Units} & \text{Value} & \text{Units} \\
(D) & 0.0249 \pm 0.0007 & \ldots & (R_p/R_*) & 0.1676 \pm 0.0025 & \ldots \\
(D) & 0.0261 \pm 0.0009 & \ldots & (R_p/R_*) & 0.1711 \pm 0.0030 & \ldots \\
(D) & 0.0276 \pm 0.0007 & \ldots & (R_p/R_*) & 0.1754 \pm 0.0024 & \ldots \\
(D) & 0.0256 \pm 0.0005 & \ldots & (R_p/R_*) & 0.1696 \pm 0.0021 & \ldots \\
(D) & 0.0237 \pm 0.0010 & \ldots & (R_p/R_*) & 0.1638 \pm 0.0034 & \ldots \\
(D) & 0.0261 \pm 0.0006 & \ldots & (R_p/R_*) & 0.1710 \pm 0.0024 & \ldots \\
(D) & 0.0259 \pm 0.0005 & \ldots & (R_p/R_*) & 0.1703 \pm 0.0021 & \ldots \\
(D) & 0.0269 \pm 0.0006 & \ldots & (R_p/R_*) & 0.1734 \pm 0.0024 & \ldots \\
(D) & 0.0274 \pm 0.0012 & \ldots & (R_p/R_*) & 0.1750 \pm 0.0038 & \ldots \\
(D) & 0.0254 \pm 0.0005 & \ldots & (R_p/R_*) & 0.1690 \pm 0.0022 & \ldots \\
(D) & 0.0277 \pm 0.0007 & \ldots & (R_p/R_*) & 0.1758 \pm 0.0025 & \ldots \\
\hline
\end{array}
\]

\[
\begin{array}{c|c|c|c|c}
\hline
\text{Other MTQ Parameters} & \text{Value} & \text{Units} & \text{Value} & \text{Units} \\
\hline
\tau_G & 0.0232 \pm 0.0007 & \text{days} & \tau_T & 0.1758 \pm 0.0025 & \text{days} \\
v_1(\nu) & 0.6767 & \ldots & v_2(\nu) & 0.1758 \pm 0.0025 & \ldots \\
\hline
\end{array}
\]

\[
\begin{array}{c|c|c|c|c}
\hline
\text{Derived Parameters} & \text{Value} & \text{Units} & \text{Value} & \text{Units} \\
\hline
b & 0.858 \pm 0.005 & \ldots & a/R_* & 5.96 \pm 0.03 & \ldots \\
\iota_{\text{orb}} & 81.72 \pm 0.09 & ^\circ \text{ (deg)} & \nu/R_* & 28.66 \pm 0.15 & \text{days}^{-1} \\
\rho_* & 2.35 \pm 0.04 & \text{g/cc} & \ldots & \ldots & \ldots \\
\hline
\end{array}
\]
Table 7

| Parameter | TMCMC | TAP | S09 | G09 | C11 |
|-----------|-------|-----|-----|-----|-----|
| $a/R_*$   | 5.97 ± 0.03 | 5.89 ± 0.05 | 5.93 ± 0.06 | ... | 6.01 ± 0.84 | ... |
| $i_{\text{orb}}$ | 81.95 ± 0.06 | 81.59 ± 0.14 | 81.85 ± 0.16 | 81.73 ± 0.13 | 81.99 ± 0.30 | ° (deg) |
| $b$       | 0.836 ± 0.003 | 0.861 ± 0.007 | 0.840 ± 0.010 | ... | 0.852 ± 0.013 | ... |
| $\rho_*$  | 2.36 ± 0.03 | 2.26 ± 0.06 | 2.30 ± 0.07 | ... | ... | g/cc |

Notes. TMCMC and TAP values are from independent analysis of APOSTLE lightcurves. S09–(Sozzetti et al. 2009), G09–(Gibson et al. 2009), C11–(Christiansen et al. 2011).

Figure 3. Plots of the joint probability distributions (JPDs) of derived system parameters from the Fixed LDC chains. Parameter estimates available in the literature are overplotted. Table 5 gives the units.

(A color version of this figure is available in the online journal.)

It is clear that using TAP on the APOSTLE data set and accounting for red-noise provides more conservative estimates of the system parameters. TAP errors are 10%–30% smaller than those reported by Sozzetti et al. (2009). Sozzetti et al. (2009) and Gibson et al. (2009) account for correlated noise by scaling the photometric errors and generally have comparable uncertainties to the APOSTLE analysis with TAP. One must note that improved system parameters derived from transit measurements, like $\rho_*$, can be used to place constraints on the age of the system (Sozzetti et al. 2009; Southworth 2010), which in turn can be used to understand the evolutionary history of the exoplanetary atmosphere. As noted by Carter & Winn (2009), we confirm that correlated noise in lightcurves is an obstacle for obtaining more precise estimates of system parameters.

6.1. Transit Depth Analysis

The auto-correlation data indicate that the Multi-Depth/Fixed LDC did not converge and hence errors for the parameters from this chain are unreliable. However, we can still study the results from the Multi-Depth fit, assuming that the best-fit values are accurate. Since the Gelman–Rubin test indicates that the parameter space was fully traversed, the best-fit points (median values) from the chain are likely to be close to values that may have resulted from a converged chain.

Figure 4 shows the transit depth $D$ as a function of transit number for $r'$-band observations of TrES-3b. The solid horizontal and dashed lines represent the best-fit value and errors, respectively, for $D$ from the Fixed LDC TMCMC fit. The dotted line is the weighted mean of transit depth values from the Multi-Depth Fixed LDC chains.
computed via TAP’s parameter set we also cannot know the level at which the depth uncertainties in the Fixed LDC chains are underestimates. In addition, there are several facts about TrES-3 that could support the notion that the depth variability seen in Figure 4 is due to stellar activity. First, the activity index reported for TrES-3 classifies it as an active star (Sozzetti et al. 2009; Knutson et al. 2010). Second, observations in the r’-band are known to be affected by spots, due to the fact that the Hz line falls in this wavelength range, and spot-to-star contrast ratios are enhanced. However, there have been no indications of spot-crossing events in the transit data (a sure sign of stellar activity); this could be due to a large impact parameter and the greater likelihood of starspots being equatorial (assuming the stellar spin-axis is aligned with the sky plane). Third, the possibility that TrES-3 is variable is evidenced by the changing flux-ratio with its comparison star (column “Flux Norm.” in Table 1); though this fact could just as easily be explained by variability in the comparison.

In summary, after accounting for incomplete and poor quality lightcurves, the resulting low-level variance in transit depth does not warrant a confident claim for the detection of variability in TrES-3. More continuous sampling of transit lightcurves without interruptions would provide better insight into variability. Moreover, as with other parameter measurements, the effect of red-noise on depth measurements needs to be studied further.

6.2. Transit Timing Analysis

Using TTVs to look for additional planets was first proposed by Agol et al. (2005) and Holman & Murray (2005): in a system where only one planet is seen in transit, a deviation from the Keplerian period could indicate the presence of additional undetected planets. TTVs on the order of minutes can be produced if an unseen companion lies close to mean motion resonance (MMR) with the transit planet (Holman et al. 2010; Lissauer et al. 2011a; Ballard et al. 2011; Nesvorný et al. 2012). We gathered transit times published by Sozzetti et al. (2009), Gibson et al. (2009), and Christiansen et al. (2011), and pooled them alongside TMCMC and TAP measurements of 11 TrES-3 transit times. The time stamps of all APOSTLE data were converted to BJD (TDB) in the customized reduction pipeline (Kundurthy et al. 2011). The APOSTLE pipeline’s time conversions have been verified by comparison to the commonly used time conversion routines made available by Eastman et al. (2010). The transit times from the literature were also converted to BJD (TDB) before comparisons were made. The Observed minus Computed (O − C) plot is shown in Figure 5. A linear ephemeris was fit to all the data using the equation

\[ T_i = T_0 + \text{Epoch}_i \times P \]

resulting in a best-fit ephemeris of

\[
T_0 = 2454185.9109932 \pm 0.0000502 \text{BJD}, \\
P = 1.306186483 \pm 0.00000070 \text{days}
\]

with a goodness of fit \( \chi^2 = 204.1 \) for 52 DOF. The large reduced chi-squared \( \chi^2/\nu = 3.92 \) indicates that a linear fit does not precisely fit the transit times, and there are significant timing deviations in the data. The largest timing deviation of \(~127 \text{ s}\) is from the Christiansen et al. (2011) data set. As previously noted, these deviations could either be due to underestimated timing errors or unaccounted timing systematics between different studies. The standard deviation of the \( O − C \) values is \(~43 \text{ s}\), hence the largest timing deviation is very close to being a 3\( \sigma \) outlier. Several features of the \( O − C \) plot seem to indicate that these timing variation are likely due to inconsistencies between the various methods used to derive transit times. For example, Christiansen et al.’s (2011) times are consistently early when compared to the expected transit times (i.e., \( O − C \)'s are all negative). Removing the Christiansen et al. (2011) transit times from the ephemeris fit results in a millisecond change in the fit period, but improves the goodness of fit slightly to \( \chi^2/\nu = 3.5 \), with the scatter being 35 s and the largest timing deviation being 78 s, a ~2\( \sigma \) offset. This fact establishes that there may be discrepancies between the methods used to derive the transit times, and hence we cannot claim any TTVs for TrES-3.

In order to compare the ephemerides derived with and without red-noise analysis, we fit for a linear ephemeris to the APOSTLE transit times from TMCMC and TAP, respectively (presented in the bottom half of Table 8). The difference between the periods derived for these subsets and the period derived from all available transit times was \(~12 \text{ ms}\). The reduced \( \chi^2 \)'s were 7.99 and 1.33 for the TMCMC and TAP subsets, respectively, confirming that TAP gives more conservative errors for the transit times thanks to the red-noise analysis. Given the much more robust fit to a linear ephemeris from TAP’s transit times, we can rule out any TTVs in the system larger than \(~27.3 \text{ s}\) (the scatter in the \( O − C \) for transit times derived from TAP). For the case when planets are in MMR, Agol et al. (2005) show that the analytic expression, \( \delta t_{\text{max}} \sim (P/4.5)j/(m_{\text{pert}} + m_{\text{trans}}) \), can roughly estimate the amplitude of the timing deviation (\( \delta t_{\text{max}} \)). The quantities \( m_{\text{pert}}, m_{\text{trans}}, P \), and \( j \) are the mass of the unseen perturber, the mass of the transiting planet, the orbital period of the transiting planet, and the order of the resonance, respectively. For the TrES-3 system, we can rule out possible system configurations given the variations seen in the \( O − C \) result from APOSTLE’s TAP fit. The maximum possible TTV amplitude that could be hidden within the variations seen in APOSTLE transit times are \( \sigma_{\text{TAP}} \sim \delta t_{\text{max}} \sim 27 \text{ s}\). Using the orbital period from Table 8 and the mass of TrES-3b, \( m_{\text{trans}} = 1.91 \text{ M}_{\oplus} \), we compute the maximum mass perturber that could exist in the TrES-3 system in the 2:1 MMR to be \(~0.66 \text{ M}_{\oplus}\), i.e., additional planets with \( M_p < 0.66 \text{ M}_{\oplus} \) may exist near the

![Figure 5. Observed minus Computed Transit Times for TrES-3b. Values from APOSTLE’s TMCMC fit, TAP and the literature are plotted. The horizontal axis represents the transit Epoch. The zero-line ephemeris is described in Section 6.2. (A color version of this figure is available in the online journal.)](image-url)
2:1 MMR. At higher order resonances, this maximum mass (for a possible perturber) is larger.

7. CONCLUSIONS

1. Photometric precision. APOSTLE monitored TrES-3b over a period of two years between 2009 and 2011, gathering 11 r′-band transit lightcurves. APOSTLE achieved photometric precision between 600 and 1200 ppm (excluding nights with poor seeing). The summary of observational results is presented in Table 1.

2. TrES-3b system parameters. From our analysis of 11 lightcurves of TrES-3b, we were able to confirm previous estimates of system parameters for the TrES-3 system. Our estimates of derived system parameters in Table 5 show improvements (~10%–30%) from previous measurements. Due to the system’s near-grazing transit, we note that several free parameters from MTQ showed strong correlations in their joint probabilities (see Figure 2). Such correlations are not suited for rapid MCMC convergence, yet we were able to produce Markov chains that converged and hence are able to derive statistically significant uncertainties for system parameters; see the results from the FLDC chain in Table 5. As previous attempts showed (Kundurthy et al. 2011), the Markov chains where limb-darkening coefficients were set as free parameters failed to converge, and so we report that APOSTLE did not achieve the photometric precision required to constrain limb-darkening coefficients for TrES-3.

3. Search for TDVs. Variations in transit depth over epoch could be evidence for stellar variability. TrES-3 is known to be variable (Sozzetti et al. 2009), hence one might have expected strong variations in the transit depth with epoch. Our uncalibrated flux ratios between the target and comparison show significant variability, yet we cannot rule out variability in the comparison star as the cause. Our Multi-Depth fits also show variations in the transit depth over transit epoch (see Figure 4 and Table 6), however, we refrain from making a confident assertion on the detection of stellar variability since (1) the parameters from the Multi-Depth chain do not have statistically significant errors as the chain failed to converge, (2) we cannot rule out inaccuracies due to the incomplete sampling of several transits, and (3) the near-grazing nature of this transit makes constraining the depth of the transit trough more challenging. In fact, removing several of the incompletely sampled transits lowers the overall scatter in transit depth measurements.

4. Search for TTVs. The transit timing precisions achieved by APOSTLE easily allow for the detection of TTV signals >1 minute (see TTV uncertainties in Table 8). This is a direct consequence of the ability of the APOSTLE program to gather lightcurves with high photometric precision from the ground. However, we were unable to detect significant timing variations for TrES-3b in our data. We find that fitting for transit parameters while accounting for red-noise provided more conservative error estimates on transit times. Transit times derived using TMCMC showed significant deviations from a linear ephemeris fit. The same data analyzed with analyzed with TAP (Carter & Winn 2009; Gazak et al. 2012) significantly reduced these deviations. The overall scatter in the O − C values from our red-noise analysis was on the order of ∼27 s, which rules out planetary companions more massive than 0.66 $M_\oplus$ near the 2:1 MMR, and larger companions near higher order resonances.

Transit times published in the literature are derived using different techniques for fitting transit parameters. A proper analysis of transit times would need a simultaneous analysis of transit lightcurves using a transit model that is (1) suited for Bayesian inference (i.e., with a fairly uncorrelated parameter set; Carter et al. 2008) and (2) a transit model that can adequately account for red-noise in the data (like TAP; Gazak et al. 2012). The catalog of transit times accumulated in the literature and used in this study of TrES-3b is not derived from such an analysis. MCMC analysis can also be inefficiently slow given large parameter sets and complex internal checks on a given model. Fast MCMC routines like The MCMC Hammer (Foreman-Mackey et al. 2012) may in the future allow for rapid analysis using complex models. We provide the entire set of APOSTLE lightcurves of TrES-3b with this publication to allow future projects to apply improved methods to study this system.

Extrapolations from Kepler planetary candidate data seem to indicate that small planets may be ubiquitous (Borucki et al. 2011). Interesting trends that have been noted are (1) hot-Jupiters tend to be alone, i.e., lacking other transiting planet siblings (Latham et al. 2011; Steffen et al. 2012) and (2) members of multi-planet systems with short period planets (period <10 days) are more likely to be hot-Neptunes (Latham et al. 764:8 (11pp), 2013 February 10, 2013).
et al. 2011; Lissauer et al. 2011b). This trend seems to reveal the existence of different formation pathways among volatile-rich planets. Hence the lack of detections TTVs in TrES-3b (a hot-Jupiter) is consistent with Kepler’s findings.

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