Outcomes of general anesthesia versus conscious sedation for Stroke undergoing endovascular treatment: a meta-analysis

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Abstract

**Background:** The impact of anesthesia strategy on the outcomes of acute ischemic stroke (AIS) patients undergoing endovascular treatment is currently controversy. Thus, we performed this meta-analysis to compare the differences of clinical and angiographic outcomes between general anesthesia (GA) and conscious sedation (CS).

**Methods:** A literature search in PubMed, Embase, and Web of Knowledge databases through February 2019 was conducted for related records on GA and CS of AIS undergoing endovascular treatment. The results of the studies were pooled and meta-analyzed with fixed- or random-effect model based on heterogeneity test in total and subgroup analyses.

**Results:** Twenty-three studies including 6703 patients were analyzed in this meta-analysis. We found that patients in the GA group have lower odds of favorable functional outcome (mRS scores ≤ 2) compared with the CS group (odds ratio [OR] = 0.62, 95% confidence interval [CI]: 0.49–0.77), and higher risk of mortality (OR = 1.68, 95% CI: 1.49–1.90), pneumonia (OR = 1.78, 95% CI: 1.40–2.26), symptomatic intracranial hemorrhage (OR = 1.64, 95% CI: 1.13–2.37). However, no significant differences were seen between the groups in the rate of recanalization (OR = 1.07, 95% CI: 0.89–1.28), vessel dissection or perforation (OR = 1.00, 95% CI: 0.98–1.03) and asymptomatic intracranial hemorrhage (OR = 1.19, 95% CI: 0.96–1.47). While in the RCT subgroup analysis, we found patients in the GA group does not show lower rate of favorable functional outcome compared with the CS group (OR = 1.84, 95% CI: 1.17–2.89). And there was no significant difference in the rate of mortality between GA and CS groups during RCT subgroup analysis (OR = 0.74, 95% CI: 0.43–1.27).

**Conclusions:** AIS patients performed endovascular treatment under GA compared with CS was associated with worse functional outcome and increased rate of mortality, but differences in worsened outcomes do not exist when one looks into the GA vs. CS RCTs. Moreover, these findings are mainly based on the retrospective studies and additional multi-center randomized controlled trials to definitively address these issues is warranted.

**Keywords:** Ischemic stroke, Anesthesia, Endovascular treatment, Meta-analysis

Background

Acute ischemic stroke (AIS) is one of the leading causes of death and long-term disability. The common therapy for AIS patients with large-vessel occlusion is endovascular treatment [1]. During the endovascular treatment, there are two types of anesthesia/sedation which are commonly used to make the AIS patients immobile, including general anesthesia (GA) and conscious sedation (CS). However, the understanding of the impact of GA or CS on the outcomes of endovascular treatment remains controversial. Previous observational studies report worse outcomes from GA than that from CS during endovascular treatment [2]. By contrast, there were some new randomized trials found that functional independence or worse tissue is either no different in the patients who had GA [3–5]. While the available previous meta-analysis studies revealed superior neurological outcome with CS compared with GA [2, 6, 7]. But those meta-analysis studies were limited by the small sample size and the
included studies were not comprehensive. Besides, there are updated and larger randomized clinical trials have conducted. In light of the continuing debate and limitations among these studies, a new and comprehensive meta-analysis study is warranted. We aim to compare the outcomes of AIS patients with GA and CS during the procedures.

Methods

Search strategy
This systematic review and meta-analysis was conducted in accordance with the PRISMA (Preferred Reporting Items for Systematic Reviews and Meta-Analyses) statement. All published articles were searched without language restricted in the PubMed, Embase, and Web of Knowledge databases through February 2019. Key words were identified and in various relevant combinations as follows: endovascular OR fibrinolytic agents OR thromboembolism OR catheter OR transcatheter OR thrombolysis OR fibrinolysis OR recanalization OR embolectomy OR thrombectomy AND intracranial embolism OR thrombosis OR stroke AND conscious sedation OR general anesthesia. The updated literature search was conducted independently by two authors (L.L., T.-F. W.). All articles were retrieved and their references were manually screened to avoid missing out other relevant articles. If data could not be extracted in the original articles, we contacted the authors to obtain them.

Study selection
The eligible studies were evaluated by two authors (T.-F.W., L.L.) independently. Disagreement was resolved by discussion, and if necessary, by a third reviewer (T.-F.W., L.L.). The inclusion criteria were as follows: (1) studies evaluating the outcomes of general anesthesia versus conscious sedation during endovascular therapy among acute ischemic stroke patients; (2) studies reporting mortality or functional outcome using the modified Rankin scale (mRS) for the general anesthesia and conscious sedation groups; and (3) the effect estimates of studies could be extracted or calculated from the available data. We excluded those studies with unavailable data to calculate or extract effect estimates. The abstracts from meeting proceedings, duplicate reports, case reports, reviews, comments, or animal studies were also excluded.

Data extraction
Two independent investigators (L.L., T.-F. W.) extracted following variables from the trials’ primary texts to ensure the reliability of the results, including the first author’s name, publication year, study period, country, inclusion criteria, exclusion criteria, outcomes, type of endovascular treatment, sample size, good outcomes (mRS scores ≤2 at 90 days) and other outcomes (including mortality, pneumonia, successful recanalization, symptomatic intracranial hemorrhage [sICH], asymptomatic intracranial hemorrhage [aICH], and vessel dissection or perforation) during the random trial. Disagreement was resolved by discussion, and if necessary, by a third reviewer (Z.-A. Z.).

Quality assessment
Quality assessment of the studies was performed by two independent reviewers (L.L., T.-F. W.). We used the Newcastle-Ottawa Scale to assess the risk of selection, comparability and exposure or outcome for case-control and cohort studies, respectively. Eight items were included to assess the quality of studies with a 9-star system. The quality score ranges from 0 to 9 stars, we judged trials as a low-quality report study while the score is 0–3 stars, while a high quality study score is at least 7 stars. And the study with 4–6 stars was defined as a moderate quality study [8]. Moreover, we used Cochrane risk of bias tool to evaluate the quality of included RCTs [9].

Statistical analysis
The outcomes in each included study, including favorable functional outcome (mRS scores ≤2 at 90 days), mortality, pneumonia, successful recanalization, sICH, aICH, and vessel dissection or perforation were extracted from primary trial results, succeeding secondary publications and their supplementary materials. For each study, odds ratios (ORs) were expressed with their 95% confidence intervals (CIs) for each outcome of interest and were calculated from patient numbers with each outcome categorized by different anesthesia type treatment. The random-effects meta-analysis (DerSimonian-Laird method) or fixed-effects meta-analysis (Mantel-Haenszel method) was used for pooling across studies and the statistical significance of pooled ORs and 95% CIs were determined with a Z test [10]. Moreover, which effects model we used was according to our heterogeneity test. To determine the degree of heterogeneity among the studies included in our meta-analysis, the I-squared ($I^2$) statistic and the Cochran Q test were used [11], with $I^2$ values less than 25% representing low heterogeneity, 25~50% representing moderate heterogeneity, and more than 75% representing high heterogeneity, respectively. When $I^2$ values was less than 50%and the P value of the Q test was more than 0.1 among the studies included in the meta-analysis, the fixed-effects model was used for pooling across studies. While the $I^2$ values was more than 50% and the P value of the Q test was less than 0.1, the random-effects model was used. Moreover, we performed sensitivity analysis by sequentially excluding each study that we have included to assess the stability of the results. To quantitate the publication bias across
included studies, the Egger regression and Begg’s methods were used [12, 13]. All statistical tests were performed with STATA software version 12.0 (StataCorp, College Station, TX, USA). Statistical significance was based on a P value < 0.05 in all analyses.

**Results**

**Characteristics of eligible studies**

The PRISMA flow diagram is shown in Fig. 1. In detail, total 2575 records were identified from PubMed, Web of Science and Embase. Of these, 1540 duplicated records and 993 unrelated records were excluded after reading the abstracts alone. Of the remaining 42 records, 19 records were excluded after further screen through full-text reading (10 records did not provide outcome of interest or available data, 6 reviews, and 3 abstracts). Finally, 23 records were eligible in our meta-analysis [3–5, 14–33], including 5 randomized controlled trials (RCTs) and 18 non-RCTs.

The characteristics of the included studies are presented in Table 1. These studies were published during 2010–2018. Ten of the 23 included studies were performed in the United States of America. The patients both in GA and CS groups have received endovascular treatment, such as IA /IV tPA, mechanical thrombectomy, stent and thromboaspiration. The baseline NIHSS scores of all patients were listed in Additional file 3: Table S1. All of the included studies were involved in the analysis of mortality rate. Majority of included studies provided information regarding the effect of GA versus CS on the incidence of mRS scores ≤2 at 90 days except for two studies [25, 28]. The results of the Newcastle-Ottawa Scale showed that 16 studies had high quality and 4 studies with a moderate quality, and no study had less than five pluses.
| Author                  | Study design                  | Study Period | Country | Inclusion Criteria | Exclusion Criteria | Outcomes                        | Type of Endovascular Treatment | Sample Size (GA/CS) | Methodological quality |
|------------------------|-------------------------------|--------------|---------|--------------------|--------------------|---------------------------------|--------------------------------|----------------------|-----------------------|
| Abou-Chebl et al.      | Retrospective cohort study    | 2005–2009    | USA     | Anterior circulation AIS | Posterior circulation strokes; no intervention was performed | mRS and mortality on day 90 | IA/IV tPA, mechanical thrombectomy, angioplasty | 428/552              | 5                     |
| Jumaa et al. (2010)   | Retrospective cohort study    | 2006–2009    | USA     | Middle cerebral artery–M1 segment occlusion treated with endovascular therapy | Verteobasilar occlusions; internal carotid artery terminus occlusion; M2 occlusion | mRS and mortality on day 90 | IA/IV tPA and mechanical thrombectomy | 53/73                | 7                     |
| Nichols et al. (2010) | Post hoc analysis of IMS II trial | NA           | USA     | Anterior circulation strokes and underwent angiography and/or intervention | Data were not available before or after angiography | Recanalization, mRS and mortality on day 90 | IA/IV tPA and mechanical thrombectomy | 26/49                | 7                     |
| Sugg et al. (2010)   | Retrospective cohort study    | 2007–2009    | USA     | AIS and underwent endovascular treatment within 8 h from symptom onset | NA | Recanalization, mRS and mortality on day 90 | Mechanical thrombectomy | 9/57                 | 5                     |
| Davis et al. (2012)  | Retrospective cohort study    | 2003–2009    | Canada  | AIS and received endovascular treatment | Management could not be determined | mRS and mortality on day 90 | IA/IV tPA and mechanical thrombectomy | 48/48                | 7                     |
| Hassan et al. (2012) | Retrospective cohort study    | 2006–2010    | USA     | AIS and received endovascular treatment | The infarct burden was greater than or equal to one third of the middle cerebral artery territory | mRS and in-hospital mortality | IV/PA, endovascular technique not specified | 53/83                | 5                     |
| Langner et al. (2013) | Retrospective cohort study    | 2005–2010    | Germany | AIS and treated with endovascular therapy | NA | mRS and mortality | IA/IV tPA and mechanical thrombectomy | 19/105               | 8                     |
| Abou-Chebl et al. (2014) | Post hoc analysis of NASA registry | NA           | North America | AIS and received endovascular treatment within 6 h from symptom onset | NA | mRS and mortality on day 90 | IV/PA, mechanical thrombectomy | 196/85               | 7                     |
| John, S et al. (2014) | Retrospective cohort study    | 2008–2012    | USA     | Anterior circulation AIS and treated with endovascular therapy | Posterior circulation AIS | In-hospital mortality and mRS on day 30 | IA/IV tPA, mechanical thrombectomy | 91/99                | 9                     |
| Li et al. (2014)      | Retrospective cohort study    | 2006–2012    | USA     | AIS and received endovascular treatment | NA | mRS and mortality | IA/IV tPA and mechanical thrombectomy | 35/74                | 8                     |
| Abou-Chebl et al. (2015) | Post hoc analysis of IMS III trial | 2006–2011    | USA     | AIS, received IV tPA within 3h and received endovascular treatment | Large regions of clear hypodensity on CT scan | mRS and in-hospital mortality | IV/PA, tPA and mechanical thrombectomy | 147/269              | 9                     |
| McDonald J.S et al. (2015) | Retrospective cohort study    | 2006–2013    | USA     | AIS and received mechanical thrombectomy | Patients who underwent another invasive surgery | In-hospital mortality and complications | IV/PA tPA and mechanical thrombectomy | 50/7                 | 7                     |
| Van Den Berg L.A. et al. (2015) | Retrospective cohort study    | 2002–2013    | Netherlands | Anterior circulation AIS | Patients lack of information | mRS and mortality | IA/IV tPA and mechanical thrombectomy | 70/278               | 7                     |
| Just, C. et al. (2016) | Retrospective cohort study    | 2000–2013    | Canada  | Underwent neuro-interventional stroke procedure | Aneurysm repair, carotid stenting, extracranial-intracranial bypass | mRS and mortality on day 90 and 180 | IA/IV tPA and mechanical thrombectomy, thromboaspiration | 42/67                | 8                     |
| Author | Study design | Study Period | Country | Inclusion Criteria | Exclusion Criteria | Outcomes | Type of Endovascular Treatment | Sample Size (GA/CS) | Methodological qualitya |
|--------|--------------|--------------|---------|-------------------|-------------------|----------|-------------------------------|-------------------|------------------------|
| Berkhemer OA. et al. (2016) | Post hoc analysis of MR CLEAN | 2010–2014 | Netherlands | AIS patients who received mechanical thrombectomy or intra-arterial thrombolysis | Cerebral hemorrhage, coagulation abnormalities | mRS and mortality on day 90 and 180 | IA tPA and mechanical thrombectomy, thromboaspiration | 79/137 | 9 |
| Schönenberger S et al. (2016) | RCT | 2014–2016 | Germany | Severe AIS, NIHSS> 10 | Diagnostic imaging results did not clearly depict site of vessel occlusion | NIHSS after 24 h, mRS and mortality on day 90 | IA tPA and angioplasty, mechanical thrombectomy, thromboaspiration | 73/77 | NA |
| Bekelis, K. et al. (2017) | Retrospective cohort study | 2009–2013 | USA | AIS patients undergoing mechanical thrombectomy | NA | Mortality in hospital | IV tPA, thromboaspiration | 441/733 | 6 |
| Lowhagen Henden, P. et al. (2017) | RCT | 2013–2016 | Sweden | Anterior circulation AIS, NIHSS>10 | Occlusion of posterior cerebral circulation | mRS and mortality on day 90 and 180 | IV tPA, mechanical thrombectomy | 45/46 | NA |
| Slezak, A. et al. (2017) | Prospective study | 2010–2015 | Switzerland | Anterior circulation AIS | NA | mRS and mortality on day 90 | IV tPA, mechanical thrombectomy | 266/135 | 7 |
| Simonsen, C. Z. et al. (2018) | RCT | 2015–2017 | Denmark | Anterior circulation AIS | Glasgow Coma Scale score <9 or premorbid mRS score >2 | mRS and mortality on day 90 and 180 | IA tPA, angioplasty, mechanical thrombectomy | 65/63 | NA |
| Peng et al. (2018) | Prospective study | 2015.01–2015.08 | China | Anterior circulation AIS | DWI lesion volume > 50 mL | Rates of successful recanalization and mRS on day 90 | Interventional treatment with Solitaire | 44/105 | 8 |
| Omer F. Eker et al. (2018) | Post hoc analysis SWIFT PRIME trial | 2013–2015 | USA and Europe | Anterior circulation AIS | Subject who is contraindicated to IV t-PA | Rates of successful recanalization and mRS on day 90 | Treated with Solitaire Revascularization Device and IV-tPA | 32/65 | 8 |
| Shanet et al. (2018) | Retrospective cohort study | 2014–2016 | China | Anterior circulation AIS | Received intra-arterial thrombolysis alone, with concomitant aneurysm or arteriovenous malformation | mRS and mortality on day 90 | NA | 114/114 | 8 |

*Newcastle Ottawa scale was designed to assess the quality of non-randomized studies. GA general anesthesia, CS conscious sedation, AIS acute ischemic stroke, IA intra-arterial, IV intra-venous, tPA tissue plasminogen activator, mRS modified Rankin Score, NA not available, RCT randomized controlled trial, NIHSS National Institute of Health Stroke Scale
on the Newcastle-Ottawa scale (Table 1). The results of quality assessment for RCTs were listed in Additional file 3: Table S2.

Outcomes

Six thousand seven hundred three patients were included in this meta-analysis in total, including 3820 patients in CS group and 2883 patients in GA group. The results of our meta-analysis suggested that GA patients have lower odds of favorable functional outcome (mRS scores ≤ 2) compared with CS patients (OR = 0.62, 95% CI: 0.49–0.77) (Fig. 2, Table 2). Moreover, GA was associated with a statistically significant higher risk of mortality (OR = 1.68, 95% CI: 1.49–1.90) (Fig. 3, Table 2), pneumonia (OR = 1.78, 95% CI: 1.40–2.26) (Table 2) and sICH (OR = 1.64, 95% CI: 1.13–2.37) (Table 2). However, there were no significant differences in the rate of recanalization (OR = 1.07, 95% CI: 0.89–1.28) (Table 2), vessel dissection or perforation (OR = 1.00, 95% CI: 0.98–1.03) (Table 2), aICH (OR = 1.19, 95% CI: 0.96–1.47) (Table 2) between the two groups.

While in the subgroup analysis, the rate of favorable functional outcome (mRS scores ≤ 2) in non-RCT subgroup also showed a statistically significant lower between CS and GA group (OR = 0.54, 95% CI: 0.44–0.66) but not in RCT subgroup (OR = 1.84, 95% CI: 1.17–2.89) (Fig. 2). Besides, GA was associated with significantly higher rate of mortality than CS in non-RCT subgroup (OR = 1.76, 95% CI: 1.55–1.99), but there was no significant difference in the rate of mortality between GA and CS groups during RCT subgroup analysis (OR = 0.74, 95% CI: 0.43–1.27) (Fig. 3).

Moreover, we have conducted an additional separate meta-analysis for high quality studies (quality scores...
Table 2 Summary of meta-analysis results

| Groups                          | Test of association | OR [95% CI] | p value | Model | Z    | Heterogeneity | \( \chi^2 \) | p value | I² (%) |
|--------------------------------|---------------------|-------------|---------|-------|------|---------------|------------|---------|--------|
| mRS score (0–2)                |                     | 0.62 [0.49–0.77] | < 0.001 | RE    | 4.16 |               | 67.83      | < 0.001 | 70.5%  |
| Mortality                      |                     | 1.68 [1.49–1.90]  | < 0.001 | FE    | 8.28 |               | 40.98      | 0.008  | 46.3%  |
| Successful recanalization      |                     | 1.07 [0.89–1.28]  | 0.943  | FE    | 0.47 |               | 26.35      | 0.023  | 46.9%  |
| Vessel dissection or perforation|                     | 1.00 [0.98–1.03]  | 0.010  | FE    | 0.19 |               | 11.21      | 0.34   | 10.8%  |
| sICH                           |                     | 1.64 [1.13–2.37]  | 0.010  | RE    | 2.59 |               | 31.63      | < 0.001 | 68.4%  |
| aICH                           |                     | 1.19 [0.96–1.47]  | 0.116  | FE    | 1.57 |               | 3.36       | 0.644  | 0.0%   |
| Pneumonia                      |                     | 1.78 [1.40–2.26]  | < 0.001 | FE    | 4.67 |               | 7.95       | 0.539  | 0.0%   |

OR odds ratio, CI confidence interval, mRS modified Rankin Score, RE random effects, FE fixed effects, sICH symptomatic intracranial hemorrhage, aICH asymptomatic intracranial hemorrhage

Fig. 3 Forest plot of meta-analysis results for the risk of mortality. OR, odds ratio, CI, confidence interval
Patients in the GA group has a higher rate of sICH than GA [34]. Moreover, in our results, we found that meta-analysis study did not demonstrate that CS was associated with poorer neurologic outcome at 90 days (OR = 0.64, 95% CI: 0.50–0.83) (Additional file 1: Figure S1) and higher mortality (OR = 1.83, 95% CI: 1.57–2.14) (Additional file 2: Figure S2) compared with CS.

**Heterogeneity and sensitivity analysis**

Obvious between-study heterogeneity ($I^2$ values more than 50%) was found for the following outcomes: good functional outcome (mRS scores ≤2) ($I^2 = 70.5\%$) and sICH ($I^2 = 68.4\%$). While no obvious heterogeneity ($I^2$ values less than 50%) was detected in death at 90 days ($I^2 = 46.3\%$), successful recanalization ($I^2 = 46.9\%$), pneumonia ($I^2 = 0.0\%$), vessel dissection or perforation ($I^2 = 0.0\%$) and aICH ($I^2 = 0.0\%$).

Each study was removed sequentially to verify the effect of each individual study in our results. There are no other important changes in our pooled OR value after excluded any study. Therefore, our results were reliable (data not shown).

**Publication bias**

Assessment of publication bias was performed by both Egger’s and Begg’s methods in our meta-analysis. And the results showed that there were no significant publication bias among the included studies (Begg’s test: $P = 0.780$, Egger’s test: $P = 0.352$).

**Discussion**

Using a comprehensive meta-analysis, we identified the worse functional outcome and higher rate of mortality among AIS patients as they received GA during endovascular treatment. Besides, we found patients in the GA group are associated with higher rate of sICH and often had more pneumonia. While no clinically meaningful differences in recanalization rate, aICH, vessel dissection or perforation were seen between patients under CS and GA. In contrast, in the RCT subgroup analysis, the difference of worse functional outcome do not exist when one looks into the GA vs. CS.

The exact reasons that why the CS group showed lower rate of mortality and better functional outcome may be multifactorial. It was well known that the purpose of general anesthesia is to decrease intraprocedural patient movement. If the patients are awake during endovascular treatment, they would move casually and be agitated during treatment, which affects Digital Subtraction Angiography images and leads to wire perforation and may result in significant vascular injury and intracranial hemorrhage. In contrast to this theory, recent meta-analysis study did not demonstrate that CS was associated with higher rate of intracranial hemorrhage than GA [34]. Moreover, in our results, we found that patients in the GA group has a higher rate of sICH. Additionally, GA has been associated with significantly higher treatment costs ($46,444 VS $30,350) [24]. And GA may limit the ability of the interventionalist to assess neurological status during the procedure. Thus, these findings do not support GA as a safer and lower cost approach for endovascular thrombectomy treatment. Moreover, as one of main factors, delay of treatment was also commonly concerned. As the post hoc analysis of MR CLEAN showed, a longer delay for patients in the GA group was observed. Intra-arterial therapy was initiated sooner after symptom onset in patients treated with non-GA as compared with GA [28]. Recently, a meta-analysis from 5 randomized controlled trials also revealed that 1 hour of delay in door-to-puncture times could reduce 19% likelihood of regaining functional independence [35]. Thus, it is reasonable to assume that treatment delays during GA may result in a disadvantage. However, the detail of door-to-puncture times among all the included studies in this meta-analysis could not be obtained completely. Therefore, future trials should study the effect of the time delay from hospital admission to vessel puncture on outcomes and its possible interaction with the type of anesthesia.

On the other hand, the poorer outcomes in GA may related with the poorer clinical status in patients who were chosen and different anesthetic agents were used for the procedure under GA. As previous studies reported, inhaled anesthetic agents were generally used for GA, which are often associated with hemodynamic disturbance, including rapid blood pressure fluctuations and lower blood pressure, which would lead to decrease of cerebral bloodflow and exacerbate ischemic injury [36, 37]. For example, Reich et al. have revealed that using propofol and the induction dose of fentanyl may cause post induction hypotension [38]. Both in the AnStroke and GOLIATH trials, blood pressure was lower in the GA group [3, 4]. In contrast, using dexmedetomidine for patients undergoing endovascular stroke treatment, which could stabilize blood pressure and prevent hypotension with induction, and improve outcomes consequently [39]. Besides, a good outcome shows an association with a higher pre-anesthesia blood pressure, while the pre-anesthesia blood pressure was negatively correlated with GA use [18, 21]. Thus, those studies indicated that the deleterious effects of GA may due to the changes of blood pressure. It is conceivable that we should pay more attention to evaluating the effects of blood pressure on outcomes and the interaction between blood pressure and the type of anesthesia should also be observed.

Different anesthetic agents may show protective or harmful effects on ischemic brain, but there are no conclusive data about the neuroprotective properties of anesthetic agents to help recommending an anesthetic agent [40].
Numerous preclinical studies indicate that isoflurane shows neuroprotective effects in ischemic preconditioning and postconditioning by alleviating glutamate excitotoxicity and opening of potassium channels [41, 42]. Besides, intravenous propofol has also been suggested as a neuroprotective agent on ischemic stroke by many molecular pathways [43]. However, these findings were just demonstrated in nonhuman primate studies. A retrospective study for endovascular management of AIS suggested that volatile anesthetics are superior to intravenous agents, but this finding should be validated by a larger randomized controlled trial [44]. Thus, to minimize the confounding effects of different drugs during endovascular treatment, the same anesthetic agent should be used as both a general anesthetic and a sedative. Among the included studies in this meta-analysis, the specific data on the type of anesthetic agent used in the GA or CS patients are unavailable in most included studies, but the GOLIATH trial has used the propofol as both a general anesthetic and a sedative in the CS group [3].

Actually, the choice of GA or CS for a given AIS patient in clinical practice, which were mainly decided by the patient’s physical status. For instance, the AIS patients with underlying medical comorbidities or stroke severity may be performed with GA as “medically indicated” [18]. Moreover, as Abou-Chebl et al. have stated that a major weakness of the retrospective study was that the association between GA and poor outcomes may be due to the AIS patients with aphasia or who were unable to follow commands and necessitated GA [24]. In non-randomized studies, the choice of GA or CS for a given AIS patient was most likely due to either technical concerns (difficulty of interventional procedure) or safety concerns (airway patency). Consequently, in the subgroup analysis of this study, we have revealed inconsistent findings between the randomized and non-randomized studies. These findings were also consistent with the previous meta-analyses by Jing R. and his colleagues, but our study have included more studies [7]. Taken together, the “medically indicated” highlights the problem of bias and may explain the reason that why the randomized and nonrandomized studies show marked discrepancy in results.

We must acknowledge that this study has several limitations. The design of included studies were various and the choice of CS or GA for a given AIS patient was not randomized for most of included studies. Thus, we conducted the subgroup analysis according to the design type of the included trials and the subgroup analysis showed inconsistent results, but we could not find the exact reasons for this discrepancy due to the lack of some of essential reported data (eg, the type and dose of anesthetic agents used, stroke location, time to treatment, baseline NIHSS scores). Although, parts of included studies have presented the baseline NIHSS scores for patients in GA and CS groups, we could not pool all the baseline NIHSS scores of patients for all included trials, because these studies presented data in different forms [mean (SD) or mean (IQR)] and the individual data were unobtainable. Moreover, in clinical practice, there was a lower rate of anterior circulation occlusions in the GA group than in the CS group. As a meta-analysis of individual patients with anterior circulation showed, outcomes were significantly better for patients who did not receive GA versus those who received GA [34]. However, we could not conduct the subgroup analyses according to those factors in the present study for relatively small or incomparable number of available studies.

**Conclusions**

In summary, the pooled data from this meta-analysis indicated that performing endovascular treatment under GA compared with CS was associated with worse functional outcome and increased rate of mortality. However, in the RCT subgroup analysis, differences in worsened outcomes do not exist between GA and CS group. Moreover, these findings are mainly based on the retrospective studies that did not randomize patients by anesthesia type. Thus, additional multi-center RCTs to definitively address these issues is warranted. In addition, to understand the exact reasons which cause the differences between the GA and CS when AIS patients are performed endovascular treatment, future studies should consider the underlying confounding factors (eg, door-to-puncture times, baseline NIHSS scores, blood pressure level).

**Additional files**

**Additional file 1:** Forest plot of meta-analysis results for good functional outcome (mRS ≤ 2) among the high quality studies. OR, odds ratio; CI, confidence interval. (TIF 3752 kb)

**Additional file 2:** Forest plot of meta-analysis results for the risk of mortality among the high quality studies. OR, Odds ratio; CI, confidence interval. (TIF 3976 kb)

**Additional file 3:** The baseline NIHSS scores of patients in each included trials. (DOCX 24 kb)

**Abbreviations**

aICH: Asymptomatic intracranial hemorrhage; AIS: Acute ischemic stroke; CI: Confidence interval; CS: Conscious sedation; FE: Fixed effects; GA: General anesthesia; IA: Intra-arterial; IV: Intra-venous; mRS: Modified Rankin scale; NIHSS: National Institute of Health Stroke Scale; OR: Odds ratio; PRISMA: Preferred Reporting Items for Systematic Reviews and Meta-Analyses; RCT: Randomized controlled trial; RE: Random effects; sICH: Symptomatic intracranial hemorrhage; tPA: Tissue plasminogen activator

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