The Efficacy of Shugan Jianpi Zhixie Therapy for Diarrhea-Predominant Irritable Bowel Syndrome: A Meta-Analysis of Randomized, Double-Blind, Placebo-Controlled Trials

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Abstract

Background
Shugan Jianpi Zhixie therapy (SJZT) has been widely used to treat diarrhea-predominant irritable bowel syndrome (IBS-D), but the results are still controversial. A meta-analysis of randomized, double-blind, placebo-controlled trials was performed to assess the efficacy and tolerability of SJZT for IBS-D.

Methods
The MEDLINE, EMBASE, Cochrane Library, the China National Knowledge Infrastructure database, the Chinese Biomedical Literature database and the Wanfang database were searched up to June 2014 with no language restrictions. Summary estimates, including 95% confidence intervals (CI), were calculated for global symptom improvement, abdominal pain improvement, and Symptom Severity Scale (BSS) score.

Results
Seven trials (N=954) were included. The overall risk of bias assessment was low. SJZT showed significant improvement for global symptom compared to placebo (RR 1.61; 95% CI 1.24, 2.10; P =0.0004; therapeutic gain = 33.0%; number needed to treat (NNT) = 3.0). SJZT was significantly more likely to reduce overall BSS score (SMD -0.67; 95% CI -0.94, -0.40; P < 0.00001) and improve abdominal pain (RR 4.34; 95% CI 2.64, 7.14; P < 0.00001) than placebo. The adverse events of SJZT were no different from those of placebo.
Conclusions

This meta-analysis suggests that SJZT is an effective and safe therapy option for patients with IBS-D. However, due to the high clinical heterogeneity and small sample size of the included trials, further standardized preparation, large-scale and rigorously designed trials are needed.

Introduction

Irritable bowel syndrome (IBS) is associated with symptoms of chronic abdominal pain, bloating and disturbed defecation in the absence of any demonstrable biochemical and anatomical abnormality [1]. About 5%-22% of general population is affected by IBS among various countries [2,3]. IBS causes significant reductions in patients’ quality of life and daily activities and involves substantial healthcare costs [4–6].

Although great progress has been made in the understanding of irritable bowel syndrome, conventional treatment remains unsatisfied. A series of systematic reviews of published drug trials were performed by the American College of Gastroenterology Task Force [7]. The quality of evidence for certain antispasmodics was graded as poor, and for tricyclic antidepressants, selective serotonin reuptake inhibitors, non-absorbable antibiotics, and C-2 chloride channel activators as moderate [7]. The trials of 5HT3 antagonists and 5HT4 agonists were of good quality, but it was noteworthy that these drugs were associated with a potential risk of ischemic colitis and cardiovascular events respectively [7].

Due to chronicity and frequency of symptoms, many patients seek alternative treatments such as traditional Chinese medicine (TCM) [8–10]. TCM is characterized by syndrome differentiation. In this regard, “stagnation of liver energy and deficiency of spleen” is considered to be the basic pathogenic factor of diarrhea-predominant IBS (IBS-D) [11]. Relieving the suppressed liver and replenishing the spleen energy for anti-diarrhea (Chinese name in pinyin “Shugan Jianpi Zhixie”) is the most important therapy in the treatment of IBS-D [12]. Tong Xie Yao Fang, an ancient formula with function of Shugan Jianpi Zhixie, has been prescribed by TCM practitioners for a long time in the treatment of IBS-D [13]. There was a systematic review of Tong Xie Yao Fang for IBS reported in 2006 [14]. In the review, twelve studies were included, but none of them were double-blind, placebo-controlled trials. The authors concluded that no definitive conclusions could be drawn due to the poor quality of the primary trials. Recently, increasing numbers of well-designed trials assessing Shugan Jianpi Zhixie therapy (SJZT) for IBS have been published [12,15–21]. However, the current state of evidence of SJZT for IBS-D has so far been unknown. Therefore, we conducted a meta-analysis of randomized, double-blind, placebo-controlled trials to determine whether SJZT is beneficial to patients with IBS-D.

Methods

Search strategy

A literature search was carried out using Medline (1989 to June 2014), EMBASE (1947 to June 2014), Cochrane Library (1993 to June 2014), the China National Knowledge Infrastructure database (1979 to June 2014), the Chinese Biomedical Literature database (1990 to June 2014) and the Wanfang database (1982 to June 2014). The search terms used were (traditional Chinese medicine OR herbal medicine OR herbs OR herbal formula OR Chinese medicinal herb joint fund (Nos. 2011B032200004 and 2012A032500004), the Guangdong Provincial Department of Science and Technology fund (Nos. 2011B031700018), China Postdoctoral Science Foundation(Nos. 2014M552187), and Administration of Traditional Chinese Medicine of Guangdong Province, China (Nos. 2012KT1145). The funders had no role in study design, data collection and analysis, decision to publish, or preparation of the manuscript.

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OR Shugan OR Jianpi OR Zhixie) AND (irritable bowel syndrome OR IBS) AND (randomized controlled trial AND (double-blind trial OR placebo-controlled trial)). No limit was placed on language. Manual searches of relevant studies supplemented the electronic searches.

Study selection
Studies meeting the following criteria were included. (i) Patients were diagnosed with IBS-D. (ii) The study was performed as a randomized, double-blind, placebo-controlled, parallel-group trial that compared the efficacy of SJZT vs. placebo. (iii) Outcomes included at least one of the following: global symptom improvement, IBS Symptom Severity Scale (BSS) score and abdominal pain improvement. Global symptom improvement was recorded as primary outcome, and overall BSS score and abdominal pain improvement were recorded as secondary outcome measures.

Data abstraction
Two researchers independently extracted data, including study design, randomization, diagnostic criteria for IBS-D, TCM criteria, sample size, dose and ingredients of each formula in the included studies, treatment duration, primary and secondary outcomes. Data were extracted as intention-to-treat analyses, in which drop-outs were assumed to be treatment failures, wherever trial reporting allowed this. Disagreements were resolved after discussion with other investigators. Assessment of methodological quality was conducted according to the Cochrane Collaboration tool.

Data synthesis and analysis
Summary relative risk (RR) and 95% confidence intervals (CI) were reported for both global symptom improvement and abdominal pain improvement. Standardized mean difference (SMD) and 95% CI were reported for BSS scores. The χ² test and the inconsistency index statistic (I²) for heterogeneity were conducted [22]. If substantial heterogeneity occurred (I² > 50% or P < 0.05), a random effect model was used to calculate the pooled RR [23]. If there was no observed heterogeneity, the pooled RR was computed by using a fixed effect model. A sensitivity analysis was done to investigate potential sources of heterogeneity between studies through omitting one trial in turn. The number needed to treat (NNT) was calculated as the reciprocal of the therapeutic gain. Begg’s test was performed to evaluate publication bias [24]. Review Manager 5.1 and Stata 12.0 were used for analyses. We calculated the optimal information size (OIS) to provide appropriate sample size for the meta-analysis [25]. To determine the OIS, a 60% control event rate for the chance of not improving symptomatic was presumed for the outcome global symptom improvement and a 25% RR reduction with a power of 80% and a two-sided α value of 0.01 [26].

Results
A total of 85 relevant studies were identified by computer search. Of these, three articles were duplicates and 43 articles were excluded on review of abstracts. After further reviewing, seven studies (N = 954) satisfied the inclusion criteria for the meta-analysis (Fig 1). A description of the included trial characteristics can be found in Table 1. The ingredients of herbal formulae were listed in Table 2.

The risk of bias assessment in the trials was summarized in Fig 2. A low risk of bias was found across studies for adequate sequence generation, allocation concealment, blinding of outcome assessment and selective outcome reporting. There were no details about placebo
content in Tang et al. study [18]. Bensoussan et al.[15] and Cai et al.[19] studies were rated as high risk for the “incomplete outcome data” item ascribing to lack of intention-to-treat (ITT) analysis.

**Primary outcome: global symptom improvement**

Among the included studies, six evaluated global symptom improvement [12,15,16–18,20]: of 917 patients, 570 were assigned to the treatment groups, whereas 347 were assigned to the placebo groups. The number of participants included in this meta-analysis was slightly larger than the calculated OIS (917 vs. 794 patients). SJZT showed significant improvement in global symptoms compared to placebo (RR 1.61; 95% CI 1.24, 2.10; P = 0.0004) (Fig 3a). 76.8% of SJZT patients had global improvement over 43.8% of placebo patients (therapeutic gain = 33.0% with NNT = 3.0) (Table 3). The heterogeneity was significant (P = 0.005, I^2 = 70.0%). We performed a sensitivity analysis to investigate potential sources of heterogeneity and found that the Leung et al. study may be the main origin of heterogeneity in the meta-analysis [12]. The heterogeneity was small after exclusion of the Leung et al. study (RR 1.79 95% CI 1.51, 2.12) (Fig 3b). The findings supported the robustness of the analysis. No evidence of asymmetry was identified by funnel plot analysis (Begg’s test P = 0.707) (Fig 4).
Secondary outcomes

**Overall BSS score.** Four of seven studies used BSS to assess the severity of IBS symptoms [15,16,18,19]. SJZT reduced the overall BSS score compared with placebo (SMD — 0.67; 95% CI — 0.94, —0.40; P < 0.00001) (Fig 5a). No substantial heterogeneity was found (P = 0.39, I² = 0.0%). Funnel plot analysis demonstrated no evidence of publication bias (Begg’s test, P = 1.000) (Fig 6).

**Abdominal pain improvement.** Two of the seven studies reported the outcome of abdominal pain improvement [16,17]. SJZT showed significant improvement for abdominal pain compared to placebo (RR 4.34; 95% CI 2.64, 7.14; P < 0.00001) (Fig 5b). No observed heterogeneity existed (P = 0.82, I² = 0.0%).
Two included studies reported daily bowel frequency [12,16], but the data could not be incorporated because Li et al. [16] reported daily bowel frequency with means and SD values, while Leung et al. [12] reported the outcome as medians.

Two trials assessed health-related quality of life—one used the validated Hong Kong Chinese version of the short form 36 (SF-36) [12], while the other used the irritable bowel

| SJZT | Ingredients of each formula |
|------|-----------------------------|
| **Bensoussan’s formula** | **Herba agastaches (Huo xiang)** | **Saposhnikovia divaricata (Fang feng)** | **Semen coicis (Yi yi ren)** | **Bupleuri chinensis (Chai hu)** |
| | **Codonopsis pilosulab (Dang shen)** | | | |
| | **Magnoliae officinalis (Hou pu)** | **Atractylodes macrocephala (Bai zhu)** | **Coptidis (Huang lian)** | **Schisandrae (Wu wei zi)** |
| | **Citri reticulatae immaturs (Qin pi)** | **Angelica dahurica (Bai zhi)** | **Poria cocos (Fu ling)** | **Plantaginis (Che qian zi)** |
| | **Paeonia alba (Bai shao)** | **Aucklandiae (Mu xiang)** | **Glycyrrhizae preparata (Zhi gan cao)** | **Coptidis (Huang lian)** |
| **Leung’s formula** | **Atractylodes macrocephala (Bai zhu)** | | | |
| | **Astragalus membranaceus (Huang qi)** | **Saposhnikovia divaricata (Fang feng)** | **Poria cocos (Fu ling)** | **Plantaginis (Che qian zi)** |
| | **Paeonia alba (Bai shao)** | **Atractylodes macrocephala (Bai zhu)** | **Coptidis (Huang lian)** | **Atractylodes chinensis (Cang zhu)** |
| | | | | |
| | **Paeonia alba (Bai shao)** | **Atractylodes macrocephala (Bai zhu)** | | |
| | **Citri reticulatae (Chen pi)** | **Murraya paniculata (Jiu li xiang)** | **Saposhnikovia divaricata (Fang feng)** | **Saposhnikovia divaricata (Fang feng)** |
| | | | | |
| | **Fructus mume (Wu mei)** | | | |
| **Li’s formula** | **Coptidis (Huang lian)** | | | |
| | **Paeonia alba (Bai shao)** | **Saposhnikovia divaricata (Fang feng)** | **Coptidis (Huang lian)** | |
| | | | | |
| | **Citri reticulatae (Chen pi)** | **Atractylodes macrocephala (Bai zhu)** | **Allii macrostemonis (Xie bai)** | |
| **Chen’s formula** | | | | |
| | **Paeonia alba (Bai shao)** | **Atractylodes macrocephala (Bai zhu)** | **Allii macrostemonis (Xie bai)** | |
| | | | | |
| | **Zingiberis preparata (Paojiang)** | **Aucklandiae (Mu xiang)** | **Saposhnikovia divaricata (Fang feng)** | **Myristicae (Rou dou kou)** |
| **Tang’s formula** | | | | |
| | **Astragalus membranaceus (Huang qi)** | **Atractylodes macrocephala (Bai zhu)** | **Coptidis (Huang lian)** | |
| | | | | |
| | **Coptidis (Huang lian)** | | | |
| **Cai’s formula** | | | | |
| | **Codonopsis pilosulab (Dang shen)** | **Atractylodes macrocephala (Bai zhu)** | **Coptidis (Huang lian)** | |
| | | | | |
| | **Glycyrrhizae preparata (Zhi gan cao)** | **Atractylodes macrocephala (Bai zhu)** | **Coptidis (Huang lian)** | |
| | | | | |
| | **Amomum vollosum (Sha ren)** | **Saposhnikovia divaricata (Fang feng)** | | |
| | | | | |
| | **Albiziae (He huan pi)** | **Saposhnikovia divaricata (Fang feng)** | | |
| | | | | |
| | **Glycyrrhizae preparata (Zhi gan cao)** | **Saposhnikovia divaricata (Fang feng)** | | |
| | | | | |
| | **Curcumae wenyujin (Yu jin)** | **Semen coicis (Yi yi ren)** | | |
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Two included studies reported daily bowel frequency [12,16], but the data could not be incorporated because Li et al. [16] reported daily bowel frequency with means and SD values, while Leung et al. [12] reported the outcome as medians.

Two trials assessed health-related quality of life—one used the validated Hong Kong Chinese version of the short form 36 (SF-36) [12], while the other used the irritable bowel

![Fig 2. Risk of bias assessment.](https://doi.org/10.1371/journal.pone.0122397.g002)
syndrome quality of life questionnaire (IBS-QOL) [18]. Both studies reported no significant difference in health-related quality of life between the SJZT group and the placebo group. Because of the heterogeneity of instruments, meta-analysis of the quality of life score could not be conducted.

Safety profile and adverse events

Safety profile was evaluated in all the included studies. Four of these reported no adverse effects during SJZT treatment [16,18–20]. Bensoussan et al. reported that two patients in the standard formula group withdrew from the study associated with gastrointestinal discomfort [15]. In Leung et al.’s study, two participants suffered from skin rash and thyroiditis respectively in the treatment group and one participant suffered from facial nerve palsy in the placebo group [12]. Chen et al. reported two cases each of mild nausea and mild pruritus [17]. The adverse events of SJZT were no different from those of placebo.

Table 3. Global symptom improvement, SJZT vs. placebo.

| Study             | Response rate, % (response/N) | Therapeutic gain, % | NNT | RR (95% CI)  |
|-------------------|-------------------------------|---------------------|-----|--------------|
|                   | SJZT                          | Placebo             |     |              |
| Bensoussan et al, 1998 | 67.4 (29/43)                     | 31.4 (11/35)                | 36.0 | 2.8          | 2.15 (1.26, 3.65) |
| Leung et al, 2006   | 35.0 (21/60)                     | 44.1 (26/59)                 | -9.1 | __           | 0.79 (0.51, 1.24) |
| Li et al, 2010      | 83.3 (25/30)                     | 60.0 (18/30)                | 23.3 | 4.3          | 1.39 (1.00, 1.94) |
| Chen et al, 2010    | 85.1 (280/329)                   | 44.2 (50/113)                | 40.9 | 2.4          | 1.92 (1.56, 2.38) |
| Tang et al, 2011    | 89.3 (25/28)                     | 36.7 (11/30)                | 52.6 | 1.9          | 2.44 (1.50, 3.96) |
| Li et al, 2014      | 72.5 (58/80)                     | 45.0 (36/80)                | 27.5 | 3.6          | 1.61 (1.22, 2.13) |
| Pooled OR          | 76.8 (438/570)                   | 43.8 (152/347)              | 33.0 | 3.0          | 1.61 (1.24, 2.10) |

IBS, irritable bowel syndrome; NNT, number needed to treat; RR, relative risk.
Discussion

The results from this meta-analysis revealed that SJZT showed significant improvement in global IBS symptoms (RR = 1.61), BSS score (SMD = -0.67), and abdominal pain (RR = 4.34) compared with placebo. No evidence of publication bias was found. Furthermore, the number of patients reporting adverse events of SJZT was similar to that of placebo.

A high placebo response rate exists in IBS patients. A series of meta-analyses showed that the proportion of placebo participants reporting global improvement was almost 40% [27–29], similar to the response effect of placebo in this meta-analysis (43.8%). The clinical benefit of SJZT for global IBS symptoms was significant, with therapeutic gains over placebo of 33.0% and NNT = 3.0. The reported NNT offered by 5-HT3 receptor antagonist was eight [30].
being said, SJZT seems to be superior to 5-HT3 receptor antagonist for IBS-D. However, randomized, double-blind trials evaluating SJZT compared with 5-HT3 receptor antagonist for IBS-D have not been found in our search.

The pathogenesis of IBS has not been fully clarified. Numerous mechanisms involve in the development of IBS, including inflammation, gut mucosal immunology, disturbed gastrointestinal motility, gut microbes, visceral hypersensitivity, altered serotonin metabolism, and psychosocial distress [31–36]. Evidence for the effectiveness of SJZT for IBS-D was identified in modern pharmacological studies. Experimental data have demonstrated that *Chang ji tai* granule can relieve diarrhea symptoms, ameliorate the stress state, inhibit bowel motility, and reduce visceral hypersensitivity in rats with IBS-D, possibly by modulating the expression of substance P mRNA in the hypothalamus and colon [37,38]. *Chang an yi hao* decoction can significantly reduce visceral hypersensitivity and enhance anti-inflammatory activities by decreasing the expression of 5-HT and regulating the levels of anti-inflammatory cytokines in rats with IBS-D [39]. Further studies in vitro and in vivo of herbal formulae need to be conducted to better understand the drug mechanism.

There was significant heterogeneity for the primary outcome. We performed a sensitivity analysis and found that the Leung et al. study may be the main origin of heterogeneity [12]. After exclusion of the Leung et al. study, the heterogeneity was effectively decreased while the corresponding pooled RR was not substantially altered. We checked all of the included studies carefully and found that there was difference of selection criteria of patients between Leung et al. study and the other selected studies. In Leung et al. study, diagnosis of IBS was based on Rome II criteria and the diarrhea-predominant type was defined by the author if diarrhea was present for at least 75% of the time during which a patient’s IBS was active [12]. The rigorous inclusion criterion of patients in Leung et al. study may contribute to the heterogeneity. Furthermore, we explored the pattern of herbal formula in Leung et al. study and found that Leung et al. added several herbs with heat-clearing functions based on SJZT, which maybe another important source of heterogeneity. Treatment based on syndrome differentiation is a characteristic of TCM. Of the included seven studies, the information for TCM syndrome classification was taken into consideration only in four studies [12,16,17,19]. For this reason, syndrome classification could be a matter of heterogeneity among the evaluated trials.
The methodological quality of included studies was general high. However, we did identified potential bias in two domains. Details about the chemical properties of placebo were not reported in Tang et al. study [18]. The extremely high response (89.3%) in the SJZT group in Tang et al. study may be related to spontaneous unblinding that the patient is aware of which treatment he or she is taking. The other flaw in the quality of included studies was the lack of ITT analysis, which could lead to incomplete outcome data. A total of 8 patients withdrew from the Bensoussan et al. study [15], of whom three patients were related to ineffective intervention. Cai et al. excluded two patients because of ineffective intervention [19]. Neither of these studies used ITT analysis and both of the studies were graded as high risk in incomplete outcome data. However, we extracted data as ITT analyses, assuming all drop-outs to be treatment failures, which could reduce the risk of attrition bias.

This meta-analysis had several limitations. First, the absolute number of clinical trials and the sample size of included studies were small. It needed to be demonstrated whether the effect size of SJZT would remain the same when applied in future large-scale trials. In the assessment of publication bias, the power of this meta-analysis was modest due to the small number of trials. Thus, there might be the possible existence of publication bias in our analysis. Second, in the included studies, patients in SJZT group were treated for 3 to 16 weeks under controlled conditions. The treatment duration was not long enough to evaluate the long term safety of SJZT for IBS-D. Third, only three of the included seven trials described follow-up evaluation [12,15,16]. Bensoussan et al. reported that, on follow-up 14 weeks after intervention completed, only the individualized formula treatment group maintained improvement rather than the standard formula group [15]. Because of the recurrent and chronic nature of IBS, long-term follow-up trials might yield different results. Therefore, it is necessary to specify the duration of the follow-up period with SJZT treatment. Fourth, discrepancy in formula composition, methods of preparation, and dose was observed between the studies, which may result heterogeneous.

In summary, the meta-analysis suggests that SJZT is more efficacious than placebo in improving global symptoms, BSS score, and abdominal pain for patients with IBS-D. SJZT is safe in short-term trials. However, due to the high clinical heterogeneity and small sample size of the included trials, further standardized preparation, randomized double-blind, multicenter, large-scale trials are required.

Supporting Information
S1 PRISMA Checklist.
(DOC)

Author Contributions
Conceived and designed the experiments: XSZ. Performed the experiments: XMS YT JRC TW FL YXK. Analyzed the data: YX YYL SHH. Contributed reagents/materials/analysis tools: YX YYL SHH. Wrote the paper: XY YYL SHH XSZ. Study supervision: XSZ RL.

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