Gestational diabetes mellitus in women born small or premature:
Systematic review and meta-analysis

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What is already known on this subject?

Approximately 10–15% of infants are born small or premature worldwide, and these children are at high risk of diseases, such as type 1 and 2 diabetes mellitus, hypertension, obesity, and kidney disease, in adulthood. A narrative review reported in 2007 that women born with LBW are at risk of gestational diabetes mellitus, but included a small number of studies. Several subsequent studies have been published since then, but there is no quantitative summary of the relevant evidence to date.

What this study adds?

This is the first systematic review and meta-analysis of observational studies that provides a comprehensive summary of evidence on the association between birth size or premature birth and future GDM risk including previously unpublished data and a large sample size. LBW, preterm birth, and SGA status may be prognostic factors for GDM.
ABSTRACT

Background: Women born preterm or with low birthweight (LBW) have an increased future risk of gestational diabetes mellitus (GDM) during pregnancy; however, a quantitative summary of evidence is lacking.

In this study, we aimed to investigate whether being born preterm, or with LBW or small for gestational age (SGA) are associated with GDM risk.

Methods: We searched the MEDLINE, Embase, and CINAHL databases and study registries, including ClinicalTrials.gov and ICTRP, from launch until 29 October 2020 for observational studies examining the association between birth weight or gestational age and GDM were eligible. We pooled the odds ratios and 95% confidence intervals using the DerSimonian and Laird random-effects model.

Results: Eighteen studies were included (N = 827,382). The meta-analysis showed that being born preterm, with LBW or SGA was associated with increased risk of GDM (pooled odds ratio = 1.84; 95% confidence interval: 1.54 to 2.20; I² = 78.3%; τ² = 0.07). Given a GDM prevalence of 2.0%, 10%, and 20%, the absolute risk differences were 1.6%, 7.0%, and 11.5%, respectively. The certainty of evidence was low due to serious concerns of risk of bias and publication bias.

Conclusion: Women born prematurely, with LBW or SGA status, may be at increased risk for GDM. However, whether this should be considered in clinical decision-making depends on the prevalence of GDM.
INTRODUCTION

Gestational diabetes mellitus (GDM) is a common pregnancy complication, with prevalence estimates being 1–36%, depending on the population studied and diagnostic criteria employed. GDM is defined as preconceptionally unconfirmed glucose intolerance identified in the second or third trimester of pregnancy. Adverse perinatal outcomes associated with uncontrolled diabetes in pregnancy include spontaneous abortion, foetal anomalies, preeclampsia, stillbirth, macrosomia, neonatal hypoglycaemia, and neonatal hyperbilirubinemia, among others. Women with a history of GDM are at a higher risk of type 2 diabetes than their counterparts.

Low birth weight (LBW) and preterm birth are the leading causes of neonatal death and childhood-onset morbidity. Approximately 10–15% of infants are born small or premature worldwide. Children who survive are at a higher risk of diseases, such as type 1 and 2 diabetes mellitus, hypertension, obesity, and kidney disease, in adulthood. The exact mechanism underlying these risks remains unclear; the Barker hypothesis proposes that pregnancy may activate biological vulnerability in utero.

A narrative review reported in 2007 that women born with LBW are at risk of GDM, but included a small number of studies, and additional research has been published subsequently. There is no quantitative summary of the relevant evidence to date. We performed the first systematic review and meta-analysis of observational studies examining the association between preterm birth, with LBW, or with SGA status and the future risk of GDM.

METHODS

We followed the Meta-analysis of Observational Studies in Epidemiology (MOOSE) guidelines (Supplementary Table S1) in the reporting of this study; the study methodology adhered to the
Evidence certainty assessment was based on the Grading of Recommendations, Assessment, Development and Evaluation (GRADE) criteria for prognostic factors. The protocol was prospectively registered with PROSPERO (CRD42020142004).

**Searches**

We searched databases such as MEDLINE, Embase, and CINAHL and study registries including ClinicalTrials.gov and ICTRP from launch until 29 October 2020. Qualified authors (YT and YK) developed the search strategy (Supplementary Table S2). No language or publication status restrictions were imposed. Reference lists of shortlisted studies were searched manually for additional potentially eligible titles.

**Study selection**

Studies were eligible for inclusion if they were observational cohort or case-control studies. Case reports or series were excluded from the present review. We included studies that involved pregnant women regardless of study setting. The exposures of interest were the infancy parameters of presently pregnant women and were defined as follows: LBW, birth weight <2500 g; small for gestational age (SGA), birth weight <10th percentile for the given gestational age, stratified by sex, using the average weight of gestational age; and preterm birth, gestational age of <37 weeks. When data on both birth weight and gestational age were reported, we extracted data on birth weight in preference. The comparator group comprised women who were not born small or born at full term.

The outcome of interest was GDM, as defined by the International Association of Diabetes Pregnancy Study Groups (IADPSG), World Health Organization (WHO), American
Diabetes Association or Endocrine Society, or International Classification of Diseases 11th
revision (ICD-11) or earlier. If studies used other definitions, they were included in the
present review; however, we removed them to assess the robustness of the pooled estimates. For
studies that reported LBW, preterm birth, or SGA as a risk factor in pregnant women without
reporting the association with GDM, we contacted study authors to acquire estimates of such
associations, where available. These additional estimates were included in the present analysis,
provided they were measures of an association between at least one of the exposure factors and
the outcome of interest.

Two investigators independently screened article titles and abstracts to shortlist relevant
studies; subsequently, the same sets of authors assessed the full text for study eligibility. In cases
where data were incomplete and precluded study eligibility assessment, we contacted study
authors with requests for clarification. Multiple publications were assessed together; the record
with the most complete data was included in the present review.

Data extraction and quality assessment

Two investigators independently extracted data from all included studies, using a pilot-tested,
uniform data extraction sheet. Any discrepancies between reviewers were resolved through
consensus between two reviewers or arbitration by a third reviewer, as required. For studies that
compared three or more exposure groups, we contacted study authors to obtain data comparing
two groups of interest. In cases where this approach was unsuitable, we extracted the relevant
data, as reported, and performed subgroup comparisons between the two groups subdivided by
specific thresholds (i.e., birth weight 2500 g, <10th percentile, and gestational age 37 weeks for
LBW, SGA, and preterm birth, respectively), as this approach may have resulted in conservative
effect estimates. The same authors who performed data extraction also independently assessed
the risk of bias in each study, using the Quality In Prognosis Studies (QUIPS) tool. We
prospectively identified the following candidate confounders: age, obesity, smoking status,
socioeconomic status, diabetes mellitus before the index gestation, and family history of diabetes.

Data synthesis and analysis
We obtained pooled and adjusted ORs with 95% CI estimates of GDM for the exposure and
control groups using the DerSimonian and Laird random-effects method. We calculated the
absolute risk difference for GDM between the exposure and control groups in low- (control
group: GDM risk was assumed to be 2.0%), moderate- (10%), and high- (20%) prevalence
groups, using the pooled odds ratios (OR) and 95% confidence intervals (CIs). This assumption
was made based on a previous report and our clinical expertise.

Publication bias was assessed qualitatively by visual inspection of the funnel plot and
quantitatively by Egger’s test. Where asymmetry was observed in the funnel plot, we
investigated the likely source of this asymmetry using the contour-enhanced funnel plot.
We evaluated between-study heterogeneity visually, using forest plots, and quantitatively, using
$I^2$ and $\tau^2$ statistics. We used the Cochrane chi-square test to calculate $I^2$ and $\tau^2$ statistics. We
performed a pre-specified subgroup analysis based on types of exposure (preterm birth, LBW, or
SGA). In pre-specified sensitivity analyses, we used crude ORs instead of adjusted ORs and
excluded studies using non-standard definitions of GDM. Some studies assessed the risk of
GDM among women born with a weight >4000 g (macrosomia); these studies were excluded.
from post-hoc sensitivity analysis, as a previous review has shown a U-shaped association between mother’s birth weight and GDM risk.  

All analyses were performed using STATA 14.2 (StataCorp LP, Texas) and RevMan 5.4 (Cochrane Collaboration, UK). Two-sided $p$-values <0.05 were considered indicative of statistical significance.

RESULTS

Figure 1 presents the flow of studies through the present review selection process. After screening 15,281 records, 59 records representing 44 studies were assessed for eligibility based on the full text. Finally, 18 studies including 827,382 participants were included in the qualitative synthesis; 15 studies including 825,622 participants were included in quantitative synthesis. Supplementary Table S3 lists all excluded studies with reasons for exclusion. We did not find any ongoing or unpublished studies by searching study registries. By contacting authors, we obtained unpublished data from two studies.  

Supplementary Table S4 shows the detail characteristics of the included studies. Nine studies (810,197 participants) used population-based samples, 2 (6,915 participants) were multicentre studies, 6 (9,439 participants) were single-centre studies, and 1 (831 participants) did not specify the study setting. Supplementary Table S5 shows the details of the inclusion and exclusion criteria of the included studies. All studies were conducted in high-income countries, mostly between the late 1990s and early 2010s. The studies included participants of non-Hispanic White, Hispanic, African, Asian, or Indian descent. Two studies (28,722 participants) only included women about to deliver their first child. Two studies (140,714 participants) compared pregnant women born preterm and at full term, 9 (216,439 participants) compared
women born with and without LBW, and 4 (468,469 participants) compared women born with
and without SGA status. The remaining 3 studies (1,760 participants) only compared the mean
birth weight of women with or without GDM. Figure S1 presents a summary of study quality
assessment using the QUIPS tool. The overall quality of the included studies was moderate to
low, mainly due to uncontrolled confounders.

Prematurity and size at birth and the risk of gestational diabetes mellitus
The median GDM rate in the control groups of the included cohort studies was 2.9% (range:
0.5% to 22%). Figure 2 presents a forest plot summarising the studies that assessed the
association between preterm birth or size at birth with GDM. Premature birth, LBW, and SGA
status were associated with a higher GDM risk (pooled OR, 1.84; 95% CI: 1.54 to 2.20; I² =
78.3%; τ² = 0.07). Supplementary Table S5 summarises the absolute risk difference in pregnant
women born with LBW, SGA status, or born preterm in the low- (2.0% risk of GDM in the
control group), medium- (10%), and high- (20%) GDM prevalence groups. The absolute risk
increases were 1.6% (95% CI: 1.0 to 2.1%), 7.0% (95% CI: 4.6 to 9.6%), and 11.5% (95% CI:
7.8 to 15.5%) in low-, moderate-, and high-prevalence settings, respectively. The certainty of
evidence was low due to serious concerns of risk of bias and publication bias.

Figure 3 presents study estimates in a funnel plot. The plot appeared asymmetrical, and
Egger’s test for funnel plot asymmetry was statistically significant (p-value = 0.030).
Supplementary Figure S2 shows the contour-enhanced funnel plot, which suggests the existence
of some missing studies on the left-hand side of the plot; these studies would have yielded
statistically non-significant findings.

Data on the birth weight of mothers with or without GDM, obtained from three studies
excluded from the meta-analysis, are presented in Supplementary Table S5. These studies
consistently reported that mothers with GDM were born with lower birth weights than those without GDM.

Subgroup and sensitivity analyses

There was substantial between-study heterogeneity ($I^2 = 78.3\%$). Figure 2 presents the results of subgroup analyses for the types of exposure (LBW, preterm birth, or SGA). Although all types of exposure were associated with GDM, there was significant heterogeneity due to the type of exposure ($p$ for interaction = 0.004). The results of additional sensitivity analyses are presented in Supplementary Figure S3, Figure S4, and Figure S5.

DISCUSSION

Main Findings

We found that women born small or premature may have future risk of GDM. However, the evidence certainty was low, and the presented findings may be overestimated, as we observed some evidence of publication bias. These findings were approximately consistent across the subgroups, including different populations, exposures, and studies of varied methodological quality; these findings were robust in sensitivity analyses.

Our finding that the mother’s size at birth or premature birth may affect GDM risk was consistent with that of a previous narrative review $^{10}$. The strength of this association was similar to that observed in women with a family history of diabetes mellitus, an established risk factor for GDM $^{29}$. However, the importance of the risk factor in clinical decision depends on the absolute risk difference. Our findings suggested that careful review of the mother’s birth status may indicate her risk of GDM and guide pregnancy management in moderate to high prevalence.
settings. The mother’s preterm birth status and size at birth are not currently considered risk factors for GDM in any of the major guidelines or risk models. Our findings may help further refine these guidelines and models or to develop new ones.

The certainty of evidence for the association between premature birth or SGA status and GDM was low due to the high risk of publication bias, as shown by funnel-plot analysis. The contour-enhanced funnel plot suggests that studies with non-significant findings may not have been published. Although we did not identify any ongoing or unpublished studies, this did not eliminate the risk of publication bias, as observational studies are less likely to be registered than clinical trials. Thus, the reported estimates may be overestimates. The studies included in this review tended not to adjust for confounders, such as smoking, obesity, socioeconomic status, and family history of diabetes. Future studies should adjust for these factors.

The main result of this review was subject to substantial between-study heterogeneity, as shown by the $I^2$ statistic. This heterogeneity may be due to the different types of exposure (LBW, SGA, or preterm birth) considered in this study. However, as all exposure types were associated with increased GDM risk, the high $I^2$ statistic may be due to the large number of participants and narrow CIs of the primary studies. Given these findings, we did not assign a low rating to the inconsistency domain of the GRADE criteria.

The underlying mechanism of the association between preterm birth or SGA status and subsequent GDM may be gestational malnutrition due to maternal malnutrition or placental insufficiency. Findings from animal studies have suggested that malnutrition in utero is associated with reduced $\beta$-cell counts, pancreas weight, and pancreatic insulin content. According to the Barker foetal origin hypothesis, these foetal programming events may affect the future risk of disease. A review of epidemiological studies has suggested that LBW and preterm
birth are associated with the risk of type 2 diabetes in adulthood; a similar mechanism is possible for GDM. 

Strengths and Limitations

A key strength of this review is that it is the first to provide a comprehensive summary of evidence on the association between birth size or premature birth and future GDM risk. This study followed the methodological recommendations presented in the Cochrane Handbook, MOOSE guidelines, and GRADE criteria. Moreover, this study included previously unpublished data and a large sample size.

Nevertheless, this study has some limitations. First, the included studies were old and may not represent the current clinical practice. The definition of GDM proposed by the IADPSG in 2010 has resulted in an increase in GDM prevalence. For example, the prevalence of GDM in the United States increased from 4.6% in 2006 to 8.2% in 2016. The median prevalence of GDM in the control groups of the included studies was 2.9%. However, empirical evidence suggests that relative effect measures are, on average, consistent across different settings; in the present study, we estimated absolute risk differences separately for low-, moderate-, and high-prevalence settings. Second, 5 of 15 studies divided birth size and preterm birth categories into three or more comparative groups, which could not be combined into two comparison groups of interest. This lack of data required methodological adjustments, as described previously. Lastly, this review only assessed certainty in estimates of association between prognostic factors and an outcome. Future studies are required to determine whether these factors can help risk-stratify pregnant women and improve the clinical management of GDM.
Conclusions

LBW, preterm birth, and SGA status may be prognostic factors for GDM. Clinicians should consider the prevalence of GDM in their setting when considering maternal preterm birth or size at birth in clinical decision-making. Due to the high likelihood of publication bias, the true association between the exposures and outcome of interest may be weaker than that reported herein. Future studies based on up-to-date diagnostic criteria, examining the dose–response relationship between exposure severity and outcome, and comparing low- and middle-income countries, are required to improve the certainty of evidence.

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COMPETING INTERESTS

The authors have no conflicts of interest to declare.

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CONTRIBUTION TO AUTHORSHIP

YT is the guarantor of the review. YT had full access to all of the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis.

Study concept and design: YT, YK, MB, ST, MK, YY

Acquisition, analysis, or interpretation of data: All authors

Drafting of the manuscript: YT, YK, MB, ST

Critical revision of the manuscript for important intellectual content: MK, YM, YY

Statistical analysis: YT

Administrative, technical, or material support: YT, YK, MB, ST

Study supervision: MK, YM, YY
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FIGURE LEGENDS

Figure 1. PRISMA flow diagram of study eligibility
Duplicate studies are displayed as a single study.

Figure 2. Risk of gestational diabetes among women born preterm, with low birth weight, or small-for-gestational-age status
Effect size (ES, represented as adjusted odds ratios); CI, confidence interval. ES was determined using the random-effects model weighted by the inverse of the variance estimate. Squares represent ES, with marker size reflecting the statistical weight of the study, obtained using
random-effects meta-analysis; horizontal lines represent 95% CIs; diamonds represent the subgroup and overall odds ratios and 95% CIs for gestational diabetes.

Figure 3. Funnel plot for the evaluation of publication bias

The solid vertical line represents the summary estimate of the association between preterm birth, low birth weight, and small for gestational age status and gestational diabetes (using random-effects meta-analysis). A significant publication bias was detected ($p = 0.030$ for Egger’s test). The funnel plot shows asymmetry, which indicates publication bias.
Records identified through database searching (MEDLINE 5109, EMBASE 11588, CINAHL 1639)

Additional records identified through other sources (n = 76)

Records after duplicates removed (n = 15281)

Records screened (n = 15281)

Records excluded (n = 15229)

Citations by handsearch (n = 8)

Full-text articles assessed for eligibility
45 studies (60 records)

Studies included in qualitative synthesis
18 studies (23 records)

Studies included in quantitative synthesis (meta-analysis)
15 studies (19 records)

Records excluded, with reasons (n=37)
- Wrong publication type (n = 1)
- Wrong populations (n = 23)
- Wrong exposure (n = 1)
- Wrong outcomes (n = 12)

Studies excluded from the quantitative synthesis due to the limited data
3 studies (4 records)
NOTE: Weights are from random effects analysis

| Study ID                  | ES (95% CI)         | Weight |
|---------------------------|---------------------|--------|
| Preterm versus normal birth |                     |        |
| a Rogvi 2012              | 1.51 (0.91, 2.52)   | 6.01   |
| Boivin 2012               | 1.20 (1.03, 1.40)   | 10.84  |
| Subtotal (I-squared = 0.0%, p = 0.392) | 1.22 (1.06, 1.42) | 16.85  |

Low versus normal birth weight

| Study ID                  | ES (95% CI)         | Weight |
|---------------------------|---------------------|--------|
| Andraweera 2019           | 2.50 (1.10, 5.68)   | 3.37   |
| Egeland 2000              | 1.80 (1.10, 2.95)   | 6.20   |
| Innes 2003                | 2.58 (1.44, 4.62)   | 5.21   |
| Ogonowski 2014            | 1.59 (1.03, 2.45)   | 6.93   |
| Olah 1996                 | 4.77 (2.56, 8.90)   | 4.82   |
| Pettit 1998               | 2.32 (0.67, 8.01)   | 1.76   |
| Savona-Ventura 2003       | 1.40 (0.79, 2.49)   | 5.29   |
| Seghieri 2002             | 1.89 (1.09, 3.28)   | 5.53   |
| Williams 1999             | 1.88 (1.56, 2.26)   | 10.44  |
| Subtotal (I-squared = 31.2%, p = 0.169) | 2.02 (1.65, 2.47) | 49.55  |

Small versus appropriate for gestational age

| Study ID                  | ES (95% CI)         | Weight |
|---------------------------|---------------------|--------|
| Chawla 2014               | 1.40 (1.30, 1.51)   | 11.54  |
| Plante 1999               | 3.60 (2.30, 5.63)   | 6.75   |
| Plante 2004               | 1.08 (0.63, 1.85)   | 5.67   |
| Laggeros 2012             | 2.30 (1.80, 2.94)   | 9.63   |
| Subtotal (I-squared = 90.3%, p = 0.000) | 1.88 (1.23, 2.89) | 33.60  |

Overall (I-squared = 78.3%, p = 0.000)

| ES (95% CI)         | Weight |
|---------------------|--------|
| 1.84 (1.54, 2.20)   | 100.00 |
Funnel plot with pseudo 95% confidence limits

- Standard error of log OR
- Log odds ratio

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