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ABSTRACT

Introduction: Volumetric biomarkers for Alzheimer disease (AD) are attractive due to their wide availability and ease of administration, but have traditionally shown lower diagnostic accuracy than measures of neuropathological contributors to AD. Our purpose was to optimize the diagnostic specificity of structural MRIs for AD using quantitative, data-driven techniques.

Methods: This retrospective study assembled several non-overlapping cohorts (total n = 1287) with publicly available data and clinical patients from Barnes-Jewish Hospital (data gathered 1990–2018). The Normal Aging Cohort (n = 383) contained amyloid biomarker negative, cognitively normal (CN) participants, and provided a basis for determining age-related atrophy in other cohorts. The Training (n = 216) and Test (n = 109) Cohorts contained participants with symptomatic AD and CN controls. Classification models were developed in the Training Cohort and compared in the Test Cohort using the receiver operating characteristics areas under curve (AUCs). Additional model comparisons were done in the Clinical Cohort (n = 579), which contained patients who were diagnosed with dementia due to various etiologies in a tertiary care outpatient memory clinic.

Results: While the Normal Aging Cohort showed regional age-related atrophy, classification models were not improved by including age as a predictor or by using volumetrics adjusted for age-related atrophy. The optimal model used multiple regions (hippocampal volume, inferior lateral ventricle volume, amygdala volume, entorhinal thickness, and inferior parietal thickness) and was able to separate AD and CN controls in the Test Cohort with an AUC of 0.961. In the Clinical Cohort, this model separated AD from non-AD diagnoses with an AUC of 0.820, an incrementally greater separation of the cohort than by hippocampal volume alone (AUC of 0.801, p = 0.06). Greatest separation was seen for AD vs. frontotemporal dementia and for AD vs. non-neurodegenerative diagnoses.

Conclusions: Volumetric biomarkers distinguished individuals with symptomatic AD from CN controls and other dementia types but were not improved by controlling for normal aging.

1. Introduction

Typical Alzheimer disease (AD)-specific biomarkers rely on in vivo detection and quantification of amyloid-β and tau, AD’s hallmark proteins. These biomarkers are increasingly used to narrow the differential diagnosis and refine the treatment of symptomatic patients in clinical practice.

Abbreviations: AD, Alzheimer disease; Adj, Adjusted for age-related atrophy; AUC, Receiver operating characteristics area under the curve; CDR, Clinical dementia rating; CI, Confidence interval; CN, Cognitively Normal; HCV, Hippocampal Volume; MMSE, Mini Mental State Exam; SARA, Select Atrophied Regions in Alzheimer disease; SUVR RSF, Standard Uptake Value Ratio (Regional spread function applied)

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practice (Rabinovici et al., 2019) based upon their appropriate use criteria (Johnson et al., 2013; Shaw et al., 2018). Despite this increased use, histological confirmation (the diagnostic reference standard) currently confirms an AD diagnosis for only 83% of AD patients at autopsy (Beach et al., 2012). In clinical trials, biomarkers can improve accuracy, utility, and cost effectiveness of screening, and can assess response to investigational therapies (Jack et al., 2018; Sevigny et al., 2016; Sperling et al., 2014). Their value is expected to further increase as AD-modifying therapies are realized, creating a need and rationale for population-level screening and an influx of patients requiring timely diagnosis and treatment (Liu et al., 2017).

Established AD biomarkers require cerebrospinal fluid analyses or positron emission tomography (PET) imaging, which are limited by expense and inaccessibility. In addition, PET imaging exposes patients to radioactivity and lumbar punctures may cause back pain, headache, and bleeding (Duits et al., 2016). These issues highlight the need for accessible, AD-specific biomarkers that can be applied to broad clinical populations. The solution may lie in brain magnetic resonance imaging (MRI), which is already standard of care in the United States for diagnostic evaluation of patients with new cognitive complaints (Knopman et al., 2001).

Current MRI biomarkers match the high accuracy of PET and cerebrospinal fluid markers in separating AD from unimpaired individuals (Frisoni et al., 2010; Morris et al., 2016). However, MRI biomarkers cannot maintain the high accuracy of amyloid biomarkers in cohorts of patients with various causes of dementia (Ossenkoppele et al., 2018; Wollman and Prohovnik, 2003). One thing likely impairing MRI-based biomarkers is the confounding influence of age-related brain atrophy or other undetected co-pathologies attributed to aging (Fotenos et al., 2005). For example, the confounding influence of preclinical AD in cohorts of cognitively normal (CN) older adults for other neuroimaging and cognitive measures has been shown previously (Brier et al., 2014; Hassenstab et al., 2016; Jack et al., 2014).

Taking this into account, we sought to optimize the diagnostic specificity of structural MRIs for AD using quantitative, data-driven techniques. Specifically, we considered whether individually adjusting volumetric measures for age-related atrophy would improve volumetric-based AD biomarkers. We compared unadjusted AD classification models to those using normal aging curves generated from CN participants free from biomarker evidence of AD. These models were validated in a research cohort with biomarker-confirmed AD and CN individuals and in a large clinical cohort containing patients with various neurodegenerative dementing diseases (including AD) who underwent MRI as part of their diagnostic evaluation for the cause of dementia.

2. Methods

2.1. Participants

The 1287 participants in the Normal Aging (n = 383), Training (n = 216), Test (n = 109), and Clinical (n = 579) Cohorts were composed from research studies or clinical patient records (collected 1990–2018). All procedures in this retrospective study were HIPAA compliant and approved by the Washington University Institutional Review Board. Informed consent was waived for the Clinical Cohort and gained for all others. Those not in the Clinical Cohort were from open source datasets and have been reported in various previous publications; the analyses of this paper and the inclusion of the participants in the Clinical Cohort are unique. All participants are described in Table 1, and Supplemental Table S1 gives these demographics separated by data source and diagnosis. All participants met the inclusion criteria described in Supplemental Table S2, which at minimum included a clinical assessment.

The Normal Aging Cohort was restricted to CN participants who had a global Clinical Dementia Rating (CDR) of 0 (Morris, 1993) stable across longitudinal follow-up, and were free of substantial AD pathology as determined by a negative amyloid PET scan (defined in Section 2.2 Imaging). See Supplemental Table S2 for full inclusion criteria and Supplemental Figure S1 for distribution of ages. These participants were sourced from Open Access Series of Imaging Studies 3 (OASIS) and the Dominantly Inherited Alzheimer Network (DIAN). OASIS is an open-source dataset that is a retrospective compilation of data for > 1000 participants. OASIS data was collected across several ongoing projects through the Knight Alzheimer Disease Research Center over the course of 30 years and includes CN controls and AD patients at various stages of impairment (LaMontagne et al., 2019). DIAN is a similar open source dataset that includes a greater number of younger controls due to its focus on dominantly inherited AD.

The Training and Test Cohorts contained CN and amyloid negative controls as well as cognitively impaired (CDR > 0) participants with a clinical diagnoses of AD and a PET scan indicating cerebral amyloidosis (full inclusion criteria in Supplemental Table S2). Participants were sourced from OASIS and the Alzheimer’s Disease Neuroimaging Initiative (ADNI). ADNI is another open source dataset that includes several hundred healthy controls and AD patients at various stages of impairment from multiple sites across the United States. Two-thirds of the CN and symptomatic AD participants from these sources (not overlapping with the controls in the Normal Aging Cohort) were randomly assigned to the Training Cohort with the remaining one-third becoming the Test Cohort. Random assignment was done separately for the CN and AD participants to maintain equal distributions of AD diagnoses.

The Clinical Cohort drew from patients seen at the Washington University Memory Diagnostic Center outpatient clinic in Saint Louis, MO. Patients were split into symptomatic AD diagnoses and various non-AD diagnoses (including CN). See Supplemental Table S2 for full inclusion criteria. Patients listed separately as “Uncertain” (153/579) did not have an etiologic cause of dementia indicated; without this, they could not be used to test the classification models. The AD and non-AD groups were also split into groups of more specific diagnoses (Supplemental Table S3).

All participants underwent a clinical assessment conducted by experienced clinicians including a semi-structured interview with the participant and a knowledgeable collateral source as well as a thorough neurological examination (see Supplemental Table S2 for full inclusion criteria). A clinical diagnoses of dementia was considered at the conclusion of each assessment, integrating results from the clinical assessment and bedside measures of cognitive function (Day et al., 2017). Dementia severity was classified using the participant’s CDR in accordance with established scoring rules (Morris, 1993). Etiologic diagnoses of dementia conformed to diagnostic criteria in use in clinical and research practices for AD (McKhann et al., 2011), dementia with Lewy bodies (McKeith et al., 2017), frontotemporal dementia (Rascovsky et al., 2011), and vascular cognitive impairment (Skrobot et al., 2018). See Supplemental Table S3 for a breakdown of specific diagnoses in the Clinical Cohort. Clinical diagnosis was made blinded to amyloid status in the OASIS participants, but not in the ADNI participants.

All CDRs and Mini Mental State Exams (MMSEs) (Folstein et al., 1975) used in this study occurred within a year of MRI, and those sourced from DIAN and ADNI all had a time difference of 0 days. Participants sourced from OASIS had an average time difference of 99 days and the Clinical Cohort had an average time difference of 118 days.

2.2. Imaging

All volumetric T1-weighted images underwent regional tissue segmentation with FreeSurfer 5.3 (freesurfer.net) using the Desikan-Killiany atlas (Desikan et al., 2006). Regional volumes (cortical and subcortical) underwent intracranial volume adjustment using a...
regression approach (Buckner et al., 2004), which fits a line to each region and the intracranial volume calculated by FreeSurfer. Studies typically fit this line to their entire cohort, we used the Normal Aging Cohort alone to mimic the conditions that would be used if the tool were to be implemented into clinical practice, enabling reproducibility at the single-subject level. Volumes after intracranial volume correction were summed across hemispheres and cortical thicknesses (not corrected for intracranial volume per standard practice) were averaged across hemispheres. For more specific imaging details see Supplemental Table S4.

Amyloid PET imaging used Florbetapir ([18F]-AV45) or [11C]-Pittsburgh compound B and was processed with an in-house pipeline using Freesurfer-derived regions (PET Unified Pipeline, github.com/ysu001/PUP) with a cerebellar cortex reference region. Partial volume correction in order to address signal spillover was done with a regional spread function (geometric transfer matrix) technique based on the scanner point spread function (determined at each imaging site) and the spread function (geometric transfer matrix) technique based on the scanner point spread function (determined at each imaging site) and the spread function (geometric transfer matrix) technique based on the scanner point spread function (determined at each imaging site) and the spread function (geometric transfer matrix).

Table 1 presents the demographic information for all cohorts. The Clinical Cohort has been separated into those given either an AD or non-AD diagnosis vs. those whose diagnosis was uncertain (and thus were not used to measure model accuracy). A ‘+’ indicates missing data: 2 MMSes, 124 Amyloids (all under age 45), and 6 Races from the Normal Aging Cohort; 1 APOE from the Training Cohort; 4 MMSes, 18 CDRs, and 40 Educations from the Clinical Cohort – Defined Diagnosis; 1 MMSE, 10 CDRs, and 11 Educations from the Clinical Cohort – Uncertain Diagnosis. Abbreviations: AD – Alzheimer disease; ADNI – Alzheimer’s Disease Neuroimaging Initiative (http://adni.loni.usc.edu); CDR – Clinical Dementia Rating; DIAN – Dominantly Inherited Alzheimer Network (https://dian.wustl.edu); MDC – Memory Diagnostic Center (http://memoryloss.wustl.edu/), soon to be available as OASIS-4; MMSE – Mini Mental State Exam; NA – Data Not Available; OASIS – Open Access Series of Imaging Studies – 3 (LaMontagne et al., 2019) (https://www.oasis-brains.org); SUVR rsf – Standard Uptake Value Ratio (Regional Spread Function applied).

2.3. Normal aging curves

To describe age-related atrophy, normal aging curves were generated for each FreeSurfer region using the Normal Aging Cohort data. For each cortical thickness and intracranial-corrected volume at each age, the mean for each cortical thickness and intracranial-corrected volume was calculated using a locally weighted scatter-plot smoother and a smoothed sliding window of two years for standard deviation. The Training, Test, and Clinical Cohorts were then adjusted for age-related atrophy by transforming the volumes and cortical thicknesses into z-scores using these age-specific means and standard deviations (Fig. 1). The hemisphere-combined volumes and thicknesses, with volumes adjusted for intracranial volume, are referred to as ‘unadjusted for age-related atrophy’. The unadjusted dataset that has undergone the z-score adjusted described above is referred to as ‘adjusted for age-related atrophy’. This adjustment for age-related atrophy greatly reduced the correlation of volumetric data with age (examples in Supplemental Table S6).

2.4. Region selection for AD classification

Volumetric measures that optimally predicted symptomatic AD relative to CN controls were selected using the Training Cohort. For 1000 iterations, a random 50% sample of the Training Cohort was fit to a least absolute shrinkage and selection operator logistic regression. All regional volumes and cortical thicknesses, as measured by FreeSurfer, participant age, sex, and scanner strength (1.5 or 3T), were included as predictors (see Supplemental Fig. S2 for entire list). This regression minimizes the sum of squared errors and has a bound on the sum of the absolute values of the coefficients, which sets many coefficients to zero. The variables not set to zero within each iteration were recorded, determining the frequency each variable was selected. This process was done using data adjusted with the normal aging curves and separately using unadjusted data (volumes still corrected by intracranial volume). The final region set included regions selected in over half the iterations for both sets of data.

2.5. Development of classification models

All classification models in this study used a logistic regression model (R package ‘stats’) fit to the Training Cohort to predict an AD diagnosis. The Age model included only chronological age as a predictor. The Hippocampal Volume (HCV) model used only hippocampal volume. The Select Atrophied Regions in Alzheimer disease (SARA) model used the regions selected in the region selection process.
described in Section 2.4. The HCVadj and SARAadj models differ from HCV and SARA in that they use data that has undergone the z-score adjustment to remove age-related atrophy, as described in Section 2.3, while the HCV and SARA models use the unadjusted data (volumes still corrected by intracranial volume). The models HCV + Age, SARA + Age, HCVadj + Age, and SARAadj + Age added chronological age as an additional predictor. In this way, the HCV adj + Age and SARAadj + Age models used age as a risk factor for AD, and separately to determine age-specific means and standard deviations when normalizing age-related changes in brain volumes and cortical thicknesses.

Each models’ receiver operating characteristics area under the curve (AUCs) were compared using the Delong method (DeLong et al., 1988) with significance set to $p < 0.003$ (Bonferroni-corrected $p < 0.05$), and confidence intervals computed using 2000 stratified bootstrap replicates. Accuracy statistics, when reported, used thresholds determined by the maximal Youden’s $J$ statistic within the Training Cohort (Youden, 1950).

3. Results

3.1. Participants

Table 1 details the demographics for each cohort, while Supplemental Table S1 breaks down demographics by data source and AD/Non-AD diagnosis.

3.2. Normal aging curves

Fig. 1 displays the age-related atrophy observed in the Normal Aging cohort, which is free of biomarker evidence of AD. Graphs of other regions can be accessed at https://github.com/benzinger-icl/SARA.

3.3. Region selection for AD classification

The hippocampal volume, inferior lateral ventricle volume, amygdala volume, entorhinal cortical thickness, and inferior parietal cortical thickness were selected in over half of the iterations in both the adjusted and unadjusted data and were thus used in all multi-region (SARA) models. Unadjusted coefficients for all models are in Table S7. Age and nucleus accumbens were additionally selected when using data adjusted for age-related atrophy, while inferior parietal volume and banks of the superior temporal sulcus volume were selected only when using unadjusted data. Frequency of selection for all regions can be found in Supplemental Fig. S2.

3.4. Classification models: impact of adjusting for age-related atrophy

AUCs for each model within the Test and Clinical Cohorts are shown in Table 2; p-values for all comparisons are in Table 3. In the Test Cohort, no significant differences were found between models using adjusted and unadjusted data. In the Clinical Cohort, HCV + Age vs HCVadj + Age, and SARA + Age vs SARAadj + Age similarly showed no statistical difference in their AUCs, but HCV and SARA had higher AUCs than their counterparts HCVadj (0.801 vs 0.743, $p < 0.001$) and SARAadj (0.820 vs 0.764, $p < 0.001$). Thus, our adjustment for age-related atrophy did not improve classification ability within our cohorts and instead lowered classification ability in models that did not include age.
3.5. Classification models: Impact of age as a predictor

In the Test Cohort, HCV and SARA models showed no statistical differences from their counterparts HCV + Age and SARA + Age (Table 3), but did outperform the Age model (0.944 vs 0.675, p < 0.001; and 0.961 vs 0.675, p < 0.001). In the Clinical Cohort, HCV and SARA had higher AUCs than their counterparts HCV + Age and SARA + Age (0.801 vs 0.792, p = 0.001 and 0.820 vs 0.799, p < 0.001), but only SARA maintained a significantly higher AUC than the Age model (0.820 vs 0.742, p = 0.002).

3.6. Classification models: Selecting a model

The AUCs of HCV and SARA were not significantly different from each other within the Test or Clinical Cohorts (Table 3), but SARA’s AUC was numerically higher than HCV’s in both cohorts (Table 2). SARA was selected as the optimal model for this reason in addition to being the only model significantly better than age alone in the Clinical Cohort. Fig. 2 provides more detail on the probabilities of AD predicted by SARA for the participants in the Test and Clinical Cohorts, as well as accuracy measures such as sensitivity and specificity. The x-axes represent the possible output from the SARA model, where a 1.0 indicates a predicted 100% probability of a symptomatic AD diagnosis. The y-axes indicate the probability density function, which is a smoothed histogram normalized to an area of 1 and allows comparison of different sized groups. An example of how to read Fig. 2a is to take the area under the curve of the CN line from x = 0 to 0.25. This is approximately 0.62, indicating that 62% of the CN Controls in the Test Cohort had a probability between 0 and 25%.

3.7. Classification models: Specific diagnoses in the clinical cohort

While models were compared using their ability to separate AD from non-AD diagnoses, the heterogeneity of the Clinical Cohort allowed us to examine more specific clinical diagnoses (groups defined in Supplemental Table S3). For non-AD diagnoses, these included other dementia types and non-dementia diagnoses that explained the cognitive complaints of the patient. For AD diagnoses, this included subgroups of AD to test how the model behaved in atypical AD patients. Fig. 3 displays the unique probability density functions for each of the more specific etiologic diagnoses in the Clinical Cohort, with the AD and non-AD diagnoses groups included for comparison. The AUCs for each sub-group are in Supplemental Table S8, but small group sizes prevented robust statistical analyses. Thus, the following is qualitative rather than an assessment of p-values.

Fig. 3a, with more detail in Supplemental Table S8, demonstrates that AD sub-groups have a high classification accuracy with only slightly lower predicted probabilities in the early-onset AD (< 65) and AD variant (such as Posterior Cortical Atrophy) groups as compared to typical (amnestic, late onset) AD. Fig. 3b indicates SARA was good at distinguishing AD from the non-neurodegenerative diagnoses, including mood and sleep disorders; in total AD vs. non-neurodegenerative diagnoses had an AUC of 0.877 (95% Confidence Interval CI): 0.833–0.922). Fig. 3c shows SARA was less able to separate AD from other neurodegenerative diagnoses and had a combined AUC of 0.719 (95% CI: 0.640–0.799). Only frontal/temporal dementia participants (subtypes combined due to small n) approached the same level of separation from AD as the non-neurodegenerative diagnoses.

The impact of using the multi-region SARA over the simple HCV model also varied by diagnosis (Supplemental Table S8). SARA had only a marginally larger AUC than HCV for separating Typical AD from non-AD diagnoses (0.827 vs 0.819), but substantial improvements were seen for separating AD variants from non-AD diagnoses (0.795 vs 0.697).

3.8. Classification models: MMSE and CDR in the clinical cohort

Fig. 4a shows the probability density plot of the Clinical Cohort separated by participants’ MMSE, while Supplemental Fig. S3a separates by CDR. MMSEs and CDRs were collected within one year of the

Table 2

| Model: | Test Cohort AUCs (95% CI) | Clinical Cohort AUCs (95% CI) |
|-------|--------------------------|-----------------------------|
| Age   | 0.675 (0.572–0.778)       | 0.742 (0.694–0.790)          |
| HCV   | 0.944 (0.902–0.987)       | 0.801 (0.756–0.846)          |
| SARA  | 0.961 (0.925–0.997)       | 0.820 (0.776–0.864)          |
| HCV + Age | 0.950 (0.909–0.991) | 0.792 (0.747–0.838) |
| SARA + Age | 0.962 (0.924–0.999) | 0.799 (0.753–0.845) |
| HCVadj | 0.948 (0.905–0.992)       | 0.743 (0.693–0.792)          |
| SARAadj | 0.952 (0.911–0.993)       | 0.764 (0.714–0.813)          |
| HCVadj + Age | 0.949 (0.908–0.991) | 0.793 (0.748–0.840) |
| SARAadj + Age | 0.961 (0.925–0.997) | 0.799 (0.752–0.845) |

Table 3

| Comparisons of Classification Models’ AUCs. | Test Cohort p-value | Clinical Cohort p-value |
|------------------------------------------|---------------------|-------------------------|
| Impact of adjusting for age-related atrophy |                     |                         |
| HCV vs HCVadj                            | 0.77 < 0.001        |                         |
| SARA vs SARAadj                          | 0.32 < 0.001        |                         |
| HCV + Age vs HCVadj + Age                | 0.45 0.48           |                         |
| SARA + Age vs SARAadj + Age              | 0.87 0.88           |                         |
| Impact of Age as a Predictor             |                     |                         |
| HCV vs HCV + Age                         | 0.20 0.001          |                         |
| SARA vs SARA + Age                       | 0.87 < 0.001        |                         |
| HCV vs Age                               | < 0.001 0.02        |                         |
| SARA vs Age                              | < 0.001 0.002       |                         |
| Single vs Multi Region Model              |                     |                         |
| HCV vs SARA                              | 0.18 0.06           |                         |

Table 3 states the p-values for the Delong tests comparing AUCs in order to select the optimal classification model. Significant differences (after accounting for multiple comparisons) are bolded. The SARA models include hippocampal volume, inferior lateral ventricle volume, entorhinal thickness, amygdala volume, and inferior parietal thickness. X + Age indicates Model X with age added as a covariate; Xadj indicates Model X using volumes and cortical thicknesses that have been adjusted for age-related atrophy.

Abbreviations: AUC – Receiver Operating Characteristics Area Under the Curve; HCV – Hippocampal Volume; SARA – Select Atrophied Regions in Alzheimer disease

Without reason to pursue the more complex processing required to adjust for age-related atrophy, further analyses were limited to unadjusted models.

| Abbreviations: AD – Alzheimer disease; AUC – Area Under the Curve; HCV – Hippocampal Volume; SARA – Select Atrophied Regions in Alzheimer disease | Test Cohort AUCs (95% CI) | Clinical Cohort AUCs (95% CI) |
|-------------------------------------------------------------------------------------------------|--------------------------|-----------------------------|
| 3.4. Classification models: Separating AD from non-AD diagnoses with MMSE and CDR |                     |                         |
| (Table 3), but did outperform the Age model (0.944 vs 0.675, p < 0.001; and 0.961 vs 0.675, p < 0.001). In the Clinical Cohort, HCV and SARA had higher AUCs than their counterparts HCV + Age and SARA + Age (0.801 vs 0.792, p = 0.001 and 0.820 vs 0.799, p < 0.001), but only SARA maintained a significantly higher AUC than the Age model (0.820 vs 0.742, p = 0.002). |
MRI (average time difference of 118 days). The strong relationship between level of impairment and SARA’s predicted probability of symptomatic AD reflects an alignment between an individual’s level of impairment and atrophy in the regions used in SARA (atrophy indicated by a model output closer to one). Despite this strong relationship, SARA maintained a fairly high AUC (0.773) within the group of participants with MMSEs 26–29 (n = 154, Fig. 4b). This indicates the SARA model had good classification ability beyond predicting level of impairment. This was similarly true for participants with a global CDR of 0.5, which had an AUC of 0.782 (n = 235, Supplemental Fig. S3b). This pattern persisted even when considering only those whose MMSE and CDR occurred within 30 days of their MRI, with AUCs of 0.806 (n = 63) and 0.771 (n = 101), respectively.

3.9. SARA as a possible clinical tool

To create an example of how SARA could be used in clinical practice, we developed multiple thresholds reflecting 80% and 90% sensitivities and specificities. Fig. 5a shows the readout a clinician might be given for an individual patient and Fig. 5b lists the percent of the clinical cohort that fell into each category. Over half of the participants who had uncertain diagnoses were given a score within the 90% sensitivity/specificity ranges for AD or non-AD diagnoses, indicating it would be a valuable tool to support clinical decision making.

4. Discussion

We demonstrated that volumetric models have excellent classification abilities that would aid in diagnosing symptomatic AD in various circumstances. We did observe region-specific atrophy even in our unique cohort of CN participants known to be without preclinical AD. However, controlling for this age-related atrophy did not improve classification. Doing so actually lowered accuracy within the Clinical Cohort if age was not included as an additional predictor. This reinforces the idea that age is a strong predictor of AD dementia and implies that these models require age or age-related atrophy to maintain the highest levels of accuracy. Thus, age-related atrophy may either convey increased risk for development of a neurodegenerative dementia illness, or, more likely, age-related atrophy may act as a proxy for age. Either way, total atrophy appears to be more predictive than atrophy specifically attributable to AD.

Our data-driven region selection approach, optimized to FreeSurfer, saw a specific pattern of atrophy in AD that overlapped with the medial temporal lobe regions reported in many previous papers. We evaluated
if using these regions would improve classification of AD. While the single region HCV and multi-region SARA models did not show statistically different AUCs, other evidence suggested SARA was the stronger classifier. First, in both the Test and Clinical Cohorts, the value of the AUC was higher in SARA than HCV. Second, in the Clinical Cohort, it was only SARA that had an AUC statistically higher than the model using age alone, without any volumetric measures. Third, the pattern of higher AUCs in SARA than in HCV was seen for most specific diagnostic group within the Clinical Cohort.

Our results suggest SARA has the greatest diagnostic specificity when distinguishing AD from frontotemporal dementia or from non-neurodegenerative diagnoses (e.g., mood disorders, sleep disorders, CN individuals). The high performance in frontotemporal dementia is especially noteworthy. FDG-PET is recommended for patients with AD vs. frontotemporal dementia diagnoses (Silverman et al., 2001), but PET scans are limited in diagnostic sensitivity and by insurance.
coverage (Medicare will cover it, but often private insurance will not),
availability, and cost. This AD vs. frontotemporal dementia differential
is often considered, especially in younger patients, and highly available
biomarkers would help identify the correct prognosis and treatment for
these patients. The Clinical Cohort had the lowest diagnostic specificity
when distinguishing AD from other neurodegenerative disorders. This
was likely due to the disorders impacting overlapping regions and pa-
tients having co-incident diagnoses. Comorbidities increase with age
and can include multiple neurodegenerative conditions, such as con-
current AD and dementia with Lewy bodies (Irwin and Hurtig, 2018).

Another way SARA reflects clinical reality is the correlation be-
tween atrophy and level of impairment measured by MMSE and global
CDR. With this in mind, we evaluated the diagnostic utility of SARA
beyond predicting impairment and found high classification ability in
the Clinical Cohort even when limited to early symptomatic partici-
pants (CDR 0.5 or MMSE 26–29). These patients are also the ones for
whom additional biomarkers would likely be most useful. These find-
ings indicate SARA is not simply acting as a proxy for MMSE or CDR,
but provides additional diagnostic information.

Fig. 4. MMSE Aligns, but is not equivalent, to predicted probability in SARA

Figure 4a displays the distributions of the SARA model’s predicted probability of AD for all participants (including Uncertain Diagnoses) in the Clinical Cohort, grouped by MMSE score instead of by diagnosis. Note the change in y-axes scale due to the tight distribution of MMSE = 30 participants. 4b displays the distribution of SARA model’s predicted probability of AD as in Fig. 2b, but only includes patients with MMSE scores of 26–29 (n = 154).

Abbreviations: AD – Alzheimer disease; AUC – Receiver Operating Characteristics Area Under the Curve; MMSE – Mini Mental State Exam; SARA – Select Atrophied Regions in Alzheimer disease
combination, we were able to demonstrate that our model approaches the sensitivity and specificity of amyloid PET and cerebrospinal fluid biomarkers. Strengths of the SARA algorithm include that it is fairly simple and transparent (unlike machine-learning algorithms), making it easier for doctors and patients to trust, will be freely available, uses MRI scans that are non-invasive and often already collected for dementia patients, and has been shown to work in both research and clinical populations. SARA and the Clinical Cohort’s data will be made available online at https://github.com/benzinger-icl/SARA and https://www.oasis-brains.org/.

While our results indicate the potential usefulness of quantitative volumetric biomarkers, there are some limitations of this study. Though our cohorts had fairly good representation of African-Americans, the general lack racial and socioeconomic diversity may bias our models. Volumetric classification may be further improved if models, including the regions used, are optimized to specific non-AD diagnoses and/or incorporate longitudinal scanning. Our use of a single set of normal aging curves and a binary AD/non-AD prediction model was due to our limited numbers, despite surpassing the sample size of many neuroimaging studies. The threshold used in the reported accuracy statistics was based upon the Training Cohort and has not been optimized to a clinical setting. This optimization would need to be validated in a separate cohort, ideally with histopathologically-confirmed diagnoses. This confirmation was not available for the Clinical Cohort and misdiagnosis may have caused an under-estimation of model accuracy. An important question for future work to address is the overlap of AD and vascular disease, which we were unable to address due to the diagnostic difficulty and limited presence of vascular problems in our research cohorts.

5. Conclusions

In summary, we have shown that an MRI-based volumetric classification model can be used to separate AD from CN controls and other dementia types. Controlling for age-related atrophy did not improve our models, and lowered performance if age was not also included as a predictor. SARA may be useful as a first step for selecting symptomatic AD participants for entrance into clinical trials or as an adjunct to the diagnostic algorithm when a clinical differential diagnosis includes AD vs. frontotemporal dementia or non-neurodegenerative conditions.

| Predicted Probability | Model Output | % of Clinical Cohort with Known Diagnosis (n=429) | % of Clinical Cohort with Uncertain Diagnosis (n=154) |
|-----------------------|--------------|------------------------------------------------|----------------------------------------------------|
| [0.0-0.29)            | <10% of those with AD | 30.8% | 37.9% |
| [0.29-0.45)           | <20% of those with AD | 8.9% | 12.4% |
| [0.45-0.61)           | Equivocal     | 13.8% | 10.4% |
| (0.61-0.82]           | <20% of those with Non-AD | 17.4% | 21.6% |
| (0.82-1.0]            | <10% of those with Non-AD | 29.1% | 17.6% |

Fig. 5. Possible Use of SARA in a Clinical Setting

Fig. 5a displays how the SARA model might be used by a clinician for a single patient, including a description of the model and multiple thresholds. The patient's specific probability of having AD as predicted by SARA is given (4%), as well as a statement reflecting that the sensitivity at that threshold is >90%, indicating both a measure of atrophy and the reliability of that measure. Fig. 5b displays the proportion of participants in the Clinical Cohort that fell into each bin of ranges of scores output from SARA. The next column shows the Clinical Cohort participants with uncertain diagnoses, with the distribution of scores suggesting that SARA would have helped provide a more certain diagnoses for the majority of the participants.

Abbreviations: AD – Alzheimer disease; SARA – Select Atrophied Regions in Alzheimer disease
Declaration of Competing Interest

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Supplementary materials

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