Recommended number of strides for automatic assessment of gait symmetry and regularity in above-knee amputees by means of accelerometry and autocorrelation analysis

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Abstract

**Background:** Symmetry and regularity of gait are essential outcomes of gait retraining programs, especially in lower-limb amputees. This study aims presenting an algorithm to automatically compute symmetry and regularity indices, and assessing the minimum number of strides for appropriate evaluation of gait symmetry and regularity through autocorrelation of acceleration signals.

**Methods:** Ten transfemoral amputees (AMP) and ten control subjects (CTRL) were studied. Subjects wore an accelerometer and were asked to walk for 70 m at their natural speed (twice). Reference values of step and stride regularity indices (Ad1 and Ad2) were obtained by autocorrelation analysis of the vertical and antero-posterior acceleration signals, excluding initial and final strides. The Ad1 and Ad2 coefficients were then computed at different stages by analyzing increasing portions of the signals (considering both the signals cleaned by initial and final strides, and the whole signals). At each stage, the difference between Ad1 and Ad2 values and the corresponding reference values were compared with the minimum detectable difference, MDD, of the index. If that difference was less than MDD, it was assumed that the portion of signal used in the analysis was of sufficient length to allow reliable estimation of the autocorrelation coefficient.

**Results:** All Ad1 and Ad2 indices were lower in AMP than in CTRL (P < 0.0001). Excluding initial and final strides from the analysis, the minimum number of strides needed for reliable computation of step symmetry and stride regularity was about 2.2 and 3.5, respectively. Analyzing the whole signals, the minimum number of strides increased to about 15 and 20, respectively.

**Conclusions:** Without the need to identify and eliminate the phases of gait initiation and termination, twenty strides can provide a reasonable amount of information to reliably estimate gait regularity in transfemoral amputees.

Background

Lower-limb amputees may present several gait deviations [1,2], with consequent increased energy cost, limited outdoor walking capacity [3] and a greater likelihood of developing muscle skeletal comorbidities [4]. In this view, several training sessions are necessary, before de-hospitalization of the patients [5-7]. However, what has been learnt during the rehabilitative/training sessions may be forgotten or wrongly recalled after some time when patients are back in their environment, and misuse of the prosthesis may occur again [3]. In our view, a “virtual gait trainer” would be very useful, i.e. an inexpensive system that the amputee can easily wear periodically to control, at home, the quality of gait in terms of symmetry and regularity and that can provide indications for improving the performances. In this context, gait symmetry is the degree of similarity of left and right steps, whereas gait regularity is the degree of...
similarity of consecutive strides. In fact, gait symmetry and regularity have been already analyzed in previous studies on subjects with lower-limb amputation [8,9].

The use of inertial sensors for analyzing gait has rapidly increased over the last few years. Recent advances in sensor manufacturing, computational power of portable systems, and wireless technology (ultimately translating into cheaper, smaller and less consuming sensors) have allowed extending the usage of inertial sensors (mainly accelerometers) beyond a merely clinical environment, to include also the home or outdoor environments, such as during activities of daily life or for long-term monitoring of signs and symptoms of specific pathologies [10-12].

In our view, key elements for designing such portable system are: 1) identification of relevant gait features and methods to measure them; 2) appropriate signal processing and feature selection; 3) evaluation of reliability and usability of the approach; 4) provision of practical guidelines for an evaluation protocol; 5) clinical validation of the evaluation protocol and of the rehabilitation outcomes. In a previous study [8] we faced the issues 1) and 2) of the aforementioned list, showing that the autocorrelation sequence of the acceleration signals measured on the thorax is appropriate to assess gait symmetry and regularity in transfemoral amputees and to provide a summary score to the user.

The present work moves from the outcomes of study [8] and aims to address items 3) and 4); in particular: i) presenting an algorithm to automatically and consistently compute symmetry and regularity indices (without the intervention of an expert operator); ii) assessing the minimum number of strides that are necessary for the appropriate assessment of gait symmetry and regularity through the autocorrelation sequence of the acceleration signals.

Methods
Participants
The subjects analyzed in this study are the same that were studied in [8]. Briefly, ten unilateral transfemoral amputees (AMP) wearing a lower-limb prosthesis with an electronically controlled knee (C-leg, Otto-Bock, D) were recruited at INAIL Prostheses Centre (Budrio, IT). All of them were confident walkers. Ten healthy subjects were also studied as a control group (CTRL). All participants were male and provided informed consent before data collection started. Main characteristics and walking parameters of the two groups of subjects are presented in Table 1.

Equipment and experimental protocol
Acceleration signals were collected by the MEMS accelerometer of an XSens inertial measurement unit (MTx, Xsens Technologies B.V., NL), which has a full scale of ± 50 m/s². The unit was placed on the thorax at the xiphoid process, following the guidelines from our previous study [8]. The sensitive axes of the accelerometer were manually aligned along the anatomical vertical (V), medio-lateral (ML), and antero-posterior (AP) axes. All the data were acquired at the sampling frequency of 100 Hz. The MTx unit applied an anti-aliasing hardware filter (1th order, cut-off frequency = 28 Hz) before digitalising the acceleration signals. Data processing and analyses were performed in Matlab (The MathWorks Inc, US).

Subjects were asked to walk straight ahead along a hallway of the INAIL Centre, for a distance of 70 m. They were initially asked to walk at their natural speed. Subsequently, they were asked to walk slower than their natural speed for a second test and then faster for a third test. The order of the tests was fixed (natural, slow, fast speed). Each subject participated in two measurement sessions: after the first three tests, the operator removed the sensor from the thorax, and the subject was asked to rest for 15 minutes; then, a second

| Table 1 Main characteristics and walking parameters of the two groups |
|-------------------|-----|-----|-----|
| **Main characteristics** | AMP | CTRL | P |
| N | 10 | 10 | - |
| Age (years) | 45.7 ± 3.1 | 27.7 ± 1.2 | < 0.0001 |
| Height (cm) | 175.9 ± 1.7 | 179.8 ± 1.5 | 0.096 |
| Body mass (Kg) | 78.5 ± 2.2 | 73.4 ± 3.1 | 0.48 |
| Prosthesis use duration (months) | 127.2 ± 38.0 | - | - |
| Electronic leg use duration (months) | 37.9 ± 10.5 | - | - |
| **General walking parameters** | | | |
| Natural walking speed (km/h) | 4.0 ± 0.2 | 4.8 ± 0.3 | 0.036 |
| CadenceV (stride/min) | 51.23 ± 0.88 | 56.59 ± 1.75 | 0.061 |
| CadenceAP (stride/min) | 51.23 ± 0.88 | 56.57 ± 1.75 | 0.062 |
| **Symmetry of step** | | | |
| Ad1V (unitless) | 0.566 ± 0.036 | 0.939 ± 0.004 | < 0.0001 |
| Ad1AP (unitless) | 0.539 ± 0.048 | 0.884 ± 0.007 | < 0.0001 |
| **Regularity of stride** | | | |
| Ad2V (unitless) | 0.016 | 0.030 | < 0.0001 |
| Ad2AP (unitless) | 0.019 | 0.038 | < 0.0001 |

AMP - amputees; CTRL - control subjects. Subscript V - Vertical; AP - Antero-posterior. Of note, cadence from V and AP signals is virtually the same, as expected. Reported values are mean ± standard error (SE).

1 with prosthesis in AMP
2 from first fitting
operator placed the sensor on the thorax, and asked the subject to repeat the three gait tests in the same order. Thus, a total of 6 gait tests were acquired for each subject, two for each gait speed.

**Gait symmetry and regularity assessed by the autocorrelation sequence**

The unbiased autocorrelation sequence of an acceleration signal \( x(i) \) can be computed by the following equation [13]:

\[
A_d(m) = \frac{1}{N - |m|} \sum_{i=1}^{N-|m|} x(i) \cdot x(i + m)
\]

in which \( N \) is the total number of samples and \( m \) is the time lag expressed as number of samples. As shown in previous studies [8,13], when the autocorrelation of the acceleration signal is computed during gait, the first peak of \( A_d(m) \), \( A_d \), reflects the regularity of the acceleration between consecutive steps of the subject. This can be interpreted as a measure of the symmetry between steps performed by the prosthetic and the sound leg (or between left and right leg in CTRL). The second peak of \( A_d(m) \), \( A_d2 \), reflects the regularity of consecutive strides. Higher \( A_d1 \) (\( A_d2 \)) values reflect higher step (stride) regularity (maximum possible value for \( A_d1 \) and \( A_d2 \) is 1). From the time lag between \( A_d1 \) and \( A_d2 \), given the sampling frequency, it is possible to compute the walking cadence (see Table 1).

Values of \( A_d1 \) computed from the acceleration signals along the vertical and antero-posterior axes will be indicated as \( A_d1V \) and \( A_d1AP \), respectively. Similar nomenclature will be used for \( A_d2 \), i.e. \( A_d2V \) and \( A_d2AP \). The medio-lateral acceleration signal, though available, was not analyzed in this study, since in our previous study [8] we found it poorly informative. Also, study [8] showed that gait tests at natural, slow and fast speeds provide essentially the same results. Thus, in this study we limited the analysis to the gait tests at the natural speed.

**Algorithm for the automatic identification of \( A_d1 \) and \( A_d2 \)**

In this study we propose a new, simple algorithm capable of automatically identifying the peaks of the autocorrelation sequence corresponding to \( A_d1 \) and \( A_d2 \) coefficients. The algorithm operates as follows: a) A search window, three samples wide, is applied to the autocorrelation sequence. The window starts from the first sample of the sequence, and is shifted sample-by-sample until the sequence ends (Figure 1). b) In each position of the search window, the maximum value of the autocorrelation sequence is considered: if it is in the second of the three samples of the window, the value and related position in the autocorrelation sequence (time shift) are saved in a vector of maxima. c) When the search of maxima is finished, the vector of maxima is analyzed; since, by definition, \( A_d2 \) shows a time shift that is twice that of \( A_d1 \), the vector is searched to identify a couple of maxima whose time shifts are in a ratio of 2, with a given tolerance (10%). If one couple is found that satisfies this criterion, and both maxima are positive, that couple is assumed as a candidate for \( A_d1 \) and \( A_d2 \) (\( A_d1 \) being the maximum with lower time shift). d) Since in some cases such couple of \( A_d1 \) and \( A_d2 \) candidates may not be the actual \( A_d1 \) and \( A_d2 \), another possible couple of maxima is searched within the vector of maxima, but now limiting the search to those maxima whose time lag does not exceed twice the time lag of the identified \( A_d2 \) candidate (within the accepted tolerance). This time interval limitation is necessary to avoid the identification of different peaks in the autocorrelation sequence (such as \( A_d3 \) or \( A_d4 \), etc.). The limitation of twice the time lag of the first \( A_d2 \) candidate was empirically found adequate. e) If another couple of \( A_d1 \) and \( A_d2 \) candidates is identified, the new \( A_d2 \) candidate value is compared to the old \( A_d2 \) candidate value: if the former is greater, the new candidates are in fact assumed as being \( A_d1 \) and \( A_d2 \); otherwise, the old candidates are assumed as being \( A_d1 \) and \( A_d2 \). This algorithm was used for the computation of \( A_d1 \) and \( A_d2 \) in all the analyses previously described. A block diagram of the algorithm is depicted in Figure 2.

**Reference values for gait symmetry and regularity**

For each of the gait tests, we initially considered only the central portion of the acceleration signals: we excluded from the analysis the first and last 1500 samples, so that we were highly confident that the transitional phases of gait initiation and termination were
completely ignored (despite these exclusions the acceleration signals were still long enough: around 4500 samples on average). The obtained \( Ad_1 \) and \( Ad_2 \) coefficient values were assumed as the reference values (\( Ad_{1\text{ref}}, Ad_{2\text{ref}} \)) for gait symmetry and regularity estimation in each test.

Gait symmetry and regularity from different portions of the acceleration signal

Firstly, we considered the central portion of the acceleration signals, as indicated above. For each signal, we performed the autocorrelation analysis for the computation of \( Ad_1 \) and \( Ad_2 \) coefficients beginning with a very small portion of the signal (100 samples). We then progressively enlarged the portion of the signal analyzed by steps of 100 samples. At each step of the procedure, \( Ad_1 \) and \( Ad_2 \) coefficients were calculated. Thus, for each acceleration signal we ended with a vector of \( Ad_1 \) and \( Ad_2 \) values that could be compared with the corresponding reference values previously obtained (see previous section: \( Ad_{1\text{ref}}, Ad_{2\text{ref}} \)).

Subsequently, we repeated the same procedure over the complete acceleration signals. Again, we obtained another (longer) vector of \( Ad_1 \) and \( Ad_2 \) values for each acceleration signal.

Standard error of measurement and minimum detectable difference

As stated above, all subjects performed each gait test twice. This allowed us to calculate the standard error of measurement, SEM [14-16]. We analyzed the \( Ad_1 \) and \( Ad_2 \) reference coefficients, obtained from the central portion of the acceleration signals. For each coefficient (\( Ad_{1\text{V}}, Ad_{1\text{AP}}, Ad_{2\text{V}}, Ad_{2\text{AP}} \)), SEM calculation was performed separately for AMP and CTRL subjects.

SEM values were used to compute the minimum detectable difference, MDD, for each \( Ad_1 \) and \( Ad_2 \) coefficient. MDD indicates the minimum difference that must be observed between two measures of one variable to assume that the difference is “real”, and not due to random errors, or systematic errors such as those, for instance, that are due to the incompetence of the operators (inter-rater variability). MDD was computed as 

\[
\text{MDD} = \text{SEM} \times 1.96 \times \sqrt{2}
\]

Minimum number of strides for proper computation of \( Ad_1 \) and \( Ad_2 \) coefficients

For each gait test, we compared \( Ad_1 \) and \( Ad_2 \) obtained by a progressively larger portion of the acceleration signal with \( Ad_{1\text{ref}}, Ad_{2\text{ref}} \). When the difference was lower than the corresponding MDD value, we assumed that the portion of the acceleration signal under analysis was sufficiently long to allow appropriate computation of the \( Ad \) coefficients. In fact, by definition, two measures of the same variable whose difference is less than the MDD have to be considered indistinguishable. From the calculated minimum length of the acceleration signal expressed in number of samples, with prior knowledge of the sampling frequency and the walking cadence of the subject (see Table 1), we were able to derive the minimum number of strides for appropriate \( Ad_1 \) and \( Ad_2 \) computation. Cadence of the subjects was computed from the reference portion of the acceleration signals.

Statistical analyses

The possible inter-rater variability in the computation of \( Ad_{1\text{ref}}, Ad_{2\text{ref}} \) was assessed through Repeated...
Measures ANOVA. To this purpose, it was used a Matlab (The MathWorks Inc, US) function, called rmaov1, which was downloaded at the URL http://www.mathworks.com/matlabcentral/fileexchange (last checked: 14 July 2011).

Repeated Measures ANOVA was also used to assess possible differences between AMP and CTRL groups in the mean value of the Ad1 and Ad2 coefficients and the main characteristics of the subjects (see Table 1). Paired t-test was performed to assess differences between couples of variables in the same group of subjects. Linear regression analysis was also performed between some variables to investigate the possible effects of subject features (see Table 1) on our results. Before any test or analysis, each variable distribution was tested for normality, and logarithmic transformation was applied in the case of non-normal distribution. As regards Ad1 and Ad2 coefficients, given their peculiar type of distribution (defined only between 0 and 1) we systematically applied the Fisher’s Z transformation. P values less than 0.05 were considered statistically significant. All the statistical analyses, except for the assessment of inter-rater variability, were performed with StatView (SAS Institute, Inc.).

Results
AMP subjects were older than CTRL, but in any case they were middle-aged (Table 1). Height and body mass were not different in the two groups. The speed of natural walking was slightly higher in CTRL, but the cadence was not significantly different in the two groups.

As regards Ad1 and Ad2 coefficients, the reference values are reported in Table 1. Both Ad1 and Ad2 coefficients were higher in CTRL than in AMP, in agreement with previous findings [8].

Reported in Table 2 are the MDD of the Ad1 and Ad2 coefficients in the two groups: MDD were computed considering and excluding the systematic effects, which are predominately ascribable to the operators. Since the operator effect was not significant (see rater P-values), MDD values were virtually identical in the two cases.

The MDD values of the Ad1 and Ad2 coefficients made it possible to perform the subsequent analyses, aimed at identifying the minimum number of strides that are necessary to obtain coefficient values not distinguishable from the reference values (i.e., different from the reference values by less than the corresponding MDD values). For sake of brevity, we present results related to Ad1AP and Ad2V only (the coefficients that in a previous study resulted in being of major interest [8]).

Figure 3 (upper panel) shows the difference (absolute values) between Ad1AP and its reference value for increasing length of the acceleration signal portion considered. Similar information is reported for Ad2V (Figure 3, lower panel). Initially, the increment in the considered acceleration signal portion determines a strong decrease of the difference, whereas subsequently the difference is small or null: when the signal portion is sufficiently long for the analysis, longer segments become useless. In contrast, when the final portion of the signal is also included, the difference can worsen due to the negative effect of the last strides (just before stopping): this is clearly observable for Ad1AP.

|               | P rater | MDD random | MDD random+rater |
|---------------|---------|------------|------------------|
| **AMP**       |         |            |                  |
| Ad1V          | 0.36    | 0.065      | 0.065            |
| Ad1AP         | 0.83    | 0.197      | 0.183            |
| Ad2V          | 0.46    | 0.098      | 0.095            |
| Ad2AP         | 0.77    | 0.095      | 0.090            |
| **CTRL**      |         |            |                  |
| Ad1V          | 0.29    | 0.023      | 0.023            |
| Ad1AP         | 0.23    | 0.068      | 0.070            |
| Ad2V          | 0.25    | 0.038      | 0.038            |
| Ad2AP         | 0.73    | 0.099      | 0.094            |

AMP - amputees; CTRL - control subjects. MDD (unitless) are reported in the case of random error only, or with inter-rater effect also considered in the analysis. P values of the inter-rater effect are also reported.
Table 3 Minimum number of strides for proper computation of Ad1{sub}AP and Ad2{sub}V

|                      | AMP             | CTRL            | P   |
|----------------------|-----------------|-----------------|-----|
| **Excluding initial and final 1500 samples in autocorrelation sequence** |                 |                 |     |
| Min. number of strides (Ad1{sub}AP) | 2.21 ± 0.23{sup 1} | 2.03 ± 0.21{sup 2} | 0.26 |
| Min. number of strides (Ad2{sub}V)  | 3.57 ± 0.44{sup 1} | 3.43 ± 0.42{sup 2} | 0.49 |
| **All samples in autocorrelation sequence** |                 |                 |     |
| Min. number of strides (Ad1{sub}AP) | 15.68 ± 2.00{sup 3} | 25.71 ± 1.85{sup 3} | 0.0014 |
| Min. number of strides (Ad2{sub}V)  | 20.81 ± 1.76{sup 4} | 33.78 ± 1.99{sup 5} | 0.0033 |

AMP - amputees; CTRL - control subjects. Two cases were considered: acceleration signals analyzed excluding initial and final portions (upper part of the table), and also including those signal portions (lower part of the table). Reported values are mean ± standard error (SE).

{sup 1} paired t-test: P = 0.0029; {sup 2} P = 0.0009; {sup 3} P = 0.0001; {sup 4} P = 0.0001

Table 3 shows the minimum number of strides required for appropriate computation of Ad1{sub}AP and Ad2{sub}V (i.e., to get values not distinguishable from the reference values). When the acceleration signals were examined excluding the transitional phases of walking the number of strides to properly compute Ad1{sub}AP was very small, and not different between AMP and CTRL subjects. The minimum number of strides for appropriate computation of Ad2{sub}V was still small but greater than that for computation of Ad1{sub}AP, for both AMP and CTRL (see paired t-test, Table 3). When the acceleration signals were examined also including the initial and final portions, the minimum number of strides for appropriate computation of Ad1{sub}AP and Ad2{sub}V largely increased, due to the negative effect produced by the strides in the transitional phases. It should be noticed that in this case the minimum number of strides was significantly lower in AMP than in CTRL.

We also examined possible relationships between the minimum number of strides for appropriate Ad1{sub}AP and Ad2{sub}V computation and the main characteristics, as well as the general walking parameters, of the subjects considered all together (see Table 1). For this analysis we considered the minimum number of strides obtained from the whole acceleration signals (see Table 3). For both Ad1{sub}AP and Ad2{sub}V, a weak inverse relationship was found with age, and a weak direct relationship with the walking speed (Figure 4), but these weak relationships seem to be mainly due to the differences between AMP and CTRL rather than being intrinsic.

Discussion

The main aim of this study was to address an issue that is crucial for the development of a system able to characterize and train gait for subjects wearing a lower-limb prosthesis, to be used out-of the hospital or rehabilitation center: the development of an automated and reliable algorithm for scores computation, and the calculation of the minimum number of strides that are necessary in these kinds of subjects for appropriate assessment of gait symmetry and regularity, through signals derived from inertial sensors and autocorrelation analysis. This would allow detecting differences in patients’ performance over time, as well as between different prosthetic prescriptions and rehabilitation strategies.

From our previous study [8], Ad1{sub}AP and Ad2{sub}V were found being the autocorrelation coefficients of major interest (best ability to estimate step symmetry and stride regularity, respectively), hence in this study we focused the analysis on these two coefficients. It should be noted that in other studies, such as [13,17], the index considered for step symmetry was Ad1/Ad2, and not Ad1. In fact, if Ad2 is low, Ad1 might be low even if there is not an actual step asymmetry. For this reason, in some studies Ad1 was normalized by Ad2. However, in our study we preferred to be conservative, and hence we considered Ad1 without any normalization.

To evaluate the minimal portion of the signal that was necessary for reliable computation of Ad1{sub}AP and Ad2{sub}V, we had to define to what extent a possible difference with the Ad1{sub}AP and Ad2{sub}V reference values could be acceptable. To this aim, we referred to the computation of SEM and MDD [14-16].

As regards the analysis excluding initial and final signal portions, our results showed that the minimum number of strides for reliable computation of Ad1{sub}AP is very small, and not different between AMP and CTRL (slightly more than 2 strides in both groups, see Table 3). This means that when the gait is in a steady-state condition (no transitional phases), the autocorrelation-based method requires just a few gait cycles for the assessment of step symmetry, and the length of signal that is needed seems not to depend on the degree of symmetry that is expected for the population examined. Indeed, the request of the method of, on average, 2 strides of signal only (see Table 3) is near to what is minimally required by construction for the autocorrelation sequence computation. Similar considerations can be reported for Ad2{sub}V. The fact that for Ad2{sub}V the minimum number of strides is higher (between 3 and 4 strides on average) is due to the reason that the
assessment of the stride quality obviously requires, by
definition, more information than that required for the
step. In fact, it is not surprising that the minimum num-
ber of strides for Ad2V is not far, on average, from being
double than for Ad1AP (see Table 3).

As regards the analysis over the whole acceleration
signal, without exclusion of the initial and final part of
the gait, the minimum number of strides for both
Ad1AP and Ad2V is, as expected, much higher than in
the previous case. The transitional phase of the gait acts
as a sort of noise in the coefficient computation, making
it more difficult to obtain coefficient values similar to
the reference values. In other words, several strides are
necessary to make the deleterious effect of the first steps
negligible. Again, the number of strides for Ad2V is lar-
ger than for Ad1AP, though the difference is less marked
than in the case with no transitional phases. Further-
more, results show that the required number of strides
for both Ad1AP and Ad2V is smaller in AMP than in
CTRL. This is due to the fact that the requirements for
reliability of coefficient computation are more severe for
CTRL (i.e., lower MDD values: see Table 2), and hence
the method, operating in a non-ideal condition (transi-
tional phases included) has more difficulties in matching
such requirements. However, in an outdoor environ-
ment, it would not be difficult to find paths of sufficient
length: even with transitional phases of gait included, a
relatively small number of strides are sufficient to make
our approach working properly (see Table 3). On the
other hand, it should be acknowledged that our results
were obtained on a plain path. If the path is not per-
fectly plain (as it could be outdoor) the recommenda-
tions about the number of strides for reliable gait
analysis may require further validation.

The statistical analysis performed to calculate SEM
and MDD values also allowed us to assess the presence
of possible inter-rater variability (every subject per-
formed the gait test twice, with the help of a different
operator in the two tests). We found that the two tests
in each subject were not different, and hence there was
no significant rater effect. Besides, this result also impli-
citly seems to suggest that test-retest variability, possibly
due to fatigue effects, was not present as well.

Our results are hardly comparable with previous results.
To our knowledge, no previous study addressed the prob-
lem of the minimum number of strides that are necessary
for reliable estimation of the quality of gait in subjects with
lower-limb prosthesis, not even with approaches different

Figure 4 Regression analysis between minimum number of strides and age or natural speed. Circle - amputees; Square - control subjects.
Left panel: minimum number of strides for Ad1AP computation (top: relationship with age; bottom: with natural walking speed). Right panel:
minimum number of strides for Ad2V computation (top: relationship with age; bottom: with natural walking speed). $R^2$ and $P$ values, and
regression equations, are reported. Variable distributions are logarithmically transformed.
from those based on inertial sensors and autocorrelation analysis. However, we can compare our results with a couple of previous studies, though not performed on amputees. In [13], Moe-Nilssen et al. seem to claim that the number of steps for adequate gait assessment through autocorrelation sequence is around ten (i.e., five strides). This is in relatively good agreement with our results obtained in ideal conditions (no transitional phases), but not with those obtained in the non-ideal case. On the other hand, in another study [18] it was reported that the suggested length of a test for reliable assessment of gait regularity was 40 m. This is indeed more similar to the results that we found in this study.

Conclusions
We addressed the problem of evaluating what the minimum number of strides is for reliable assessment of gait symmetry and regularity through accelerometry-based autocorrelation analysis in subjects with lower-limb prosthesis. We were able to reach the following conclusions and recommendations: when the gait includes transitional phases due to gait initiation, the number of strides to be performed over a rectilinear path should be around 15 for the assessment of step symmetry and 20 for stride regularity.

List of abbreviations
Adr: t-th peak of the autocorrelation sequence; ANOVA: Analysis of variance; AMP: Subject with amputation; AP: Antero-posterior axis; CTRL: Control subject; MDD: Minimum detectable difference; ML: Medio-lateral axis; Ref: Reference; V: Vertical axis; SEM: Standard error of measurement.

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Authors` contributions
AT has made substantial contributions to analysis and interpretation of data and has been involved in drafting the manuscript. LR has made substantial contributions to analysis and interpretation of data and has been involved in revising the manuscript. MR has made substantial contributions to acquisition, analysis and interpretation of data. AGC has made substantial contributions to conception and design, analysis and interpretation of data, and has been involved in revising the manuscript. LC has made substantial contributions to conception and design of the study and has been involved in revising the manuscript. All authors read and approved the final manuscript.

Competing interests
The authors declare that they have no competing interests.

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References
1. Segal AD, Orenduff MS, Klute GK, McDowell ML, Pecoraro JA, Shaffer J, Czernecki JM. Kinematic and kinetic comparisons of transfemoral amputee gait using C-Leg and Mach SNS prosthetic knees. J Rehabil Res Dev 2006, 43:657-670.
2. Hillman SJ, Donald SC, Herman J, McCurrach E, McGarry A, Richardson AM, Robb JE. Repeatability of a new observational gait score for unilateral lower limb amputees. Gait Posture 2010, 32:19-45.
3. Hagberg K, Hagstrom E, Brannemark R. Physiological cost index (PCI) and walking performance in individuals with transfemoral prostheses compared to healthy controls. Disabil Rehabil 2007, 29:643-649.
4. Galey R, Allen K, Castles J, Kucharka J, Roeder M. Review of secondary physical conditions associated with lower-limb amputation and long-term prosthetic use. J Rehabil Res Dev 2008, 45:15-29.
5. Christensen B, Elegard B, Breiter U, Ostrup E. The effect of prosthetic rehabilitation in lower limb amputees. Prosthet Orthot Int 1995, 19:46-52.
6. Summers GD, Morrison JD, Cochrane GM. Amputee walking training: a preliminary study of biomechanical measurements of stance and balance. Int Disabil Stud 1988, 10:1-5.
7. Gauthier-Gagnon C, Grise MC, Potvin D. Enabling factors related to prosthetic use by people with transfibial and transfemoral amputation. Arch Phys Med Rehabil 1999, 80:706-713.
8. Tura A, Raggi M, Rocchi L, Cutti AG, Chiai L. Gait symmetry and regularity in transfemoral amputees assessed by trunk accelerations. J Neuroeng Rehabil 2010, 7:4.
9. Nolan L, Witt A, Duddzinski K, Lees A, Lake M, Wychowanski M. Adjustments in gait symmetry with walking speed in trans-femoral and trans-tibial amputees. Gait Posture 2003, 17:142-151.
10. Mathie MJ, Coster AC, Lowell NH, Celler BG. Accelerometry: providing an integrated, practical method for long-term, ambulatory monitoring of human movement, Physiol Meas 2004, 25:R1-20.
11. Bonato P. Clinical applications of wearable technology, Conf Proc IEEE Eng Med Biol Soc 2004, 2009:6580-6583.
12. Patel S, Lorincz K, Hughes R, Huggins N, Crowdon J, Standaert D, Akay M, Dy J, Welsh M, Bonato P. Monitoring motor fluctuations in patients with Parkinson’s disease using wearable sensors. IEEE Trans Int Technol Biomed 2009, 13:864-873.
13. Moe-Nilssen R, Helbostad JL. Estimation of gait cycle characteristics by trunk accelerometry. J Biomech 2004, 37:121-126.
14. Campanini I, Merlo A. Reliability, smallest real difference and concurrent validity of indices computed from GRF components in gait of stroke patients. Gait Posture 2009, 30:127-131.
15. de Vet HC, Terwee CB, Knol DL, Bouter LM. When to use agreement versus reliability measures. J Clin Epidemiol 2006, 59:1033-1039.
16. Weir JP. Quantifying test-retest reliability using the intraclass correlation coefficient and the SEM. J Strength Cond Res 2005, 19:231-240.
17. Tenzer P, Devia O, Melchta A, Luthi F. Prescription footwear for severe injuries of foot and ankle: effect on regularity and symmetry of the gait assessed by trunk accelerometry. Gait Posture 2009, 30:492-496.
18. Auvinet B, Chaleil D, Barrey E. Accelerometric gait analysis for use in hospital outpatients. Rev Rhum Engl Ed 1999, 66:589-597.