The nexus between health status and health expenditure, energy consumption and environmental pollution: Empirical evidence from SAARC-BIMSTEC regions

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Abstract

Background: The prevailing COVID-19 pandemic made us deep concern for considering the betterment of health status along with health expenditure, energy, and environmental issues all over the world. This paper examines the nexus between health status and health expenditures (both public and private), energy consumption and environmental pollution in the SAARC-BIMSTEC region.

Methodology: We have utilized the panel autoregressive distributed lag (ARDL) model, heterogeneous panel causality test, cross sectional dependence test, cointegration test and Pesaran cross sectional dependent (CADF) unit root test for obtaining estimated results from the data of 16 years (2002-2017).

Results: The acquired results authorize the cointegration among the variables used, where energy consumption, public and private health expenditure and economic growth have positive and statistically significant effects and environmental pollution has both negative and significant effects on the health status of these regions in the long-run, but no panel wise significant impact is found in the short-run. Two-way causal relationships between health status and environmental pollution, public and private health expenditure, economic growth and sanitation facilities, and a one-way causality running from energy consumption to health status are revealed.

Conclusions: The improved health status in the SAARC-BIMSTEC region required to be protected by articulating the effective policies on both public and private health expenditures, environmental pollution, energy consumption, and economic growth. The attained results are theoretically and empirically consistent, and have important policy implications in the health sector.

Keywords: Health status; energy consumption; environmental pollution; health expenditure; panel ARDL.

JEL codes: I10, I15, I18, C23
1. Introduction

Health is pondered to be one of the fundamental rights of human beings but the recent (COVID-19) and previous pandemics, along with numerous epidemics have made the deeper concern for re-thinking a number of health related issues across the globe. The world is now experiencing a huge toll on health as well as a massive death toll due to ignorance and lack of proper consideration of the various health related factors like energy, environment, public and private health expenditures, hygiene and sanitation. In many areas, health facilities have been found to be wanting, exacerbating the health related issues. Therefore, ensuring better health facilities for all people in order to build a safer world has now become a prime policy goal across the globe. In 2015 the United Nations (UN) had proposed the Sustainable Development Goals (SDGs) to be achieved by 2030, citing good health and well-being as among its principal targets, incorporating affordable and clean energy, pure water and sanitation to improve the health of people, no matter where they live (Goal-3, 6 and 7, UN, 2015; WHO, 2018). The challenges resulting from COVID-19 have made the implementation of these goals imperative.

In the SAARC-BIMSTEC region, the notion of health status, health expenditures, energy use, environment and economic growth are essential considerations that require prioritization. These regions comprise the ten emerging countries\(^1\) covering the geographical areas of 5,935,573.999 square kilometres with a population is 1959.45 million, many of whom are living in poverty (average rate 20.54% approximately), where awareness of health issues is small and therefore requires enough attention regarding this (WDI, 2020). In 2017, the average life expectancy at birth in this region was 71.18 years, where the highest belonged to Thailand (76.68 years) and the lowest belonged to Afghanistan (64.13 years) (WDI, 2020). Following this, the average public and private health expenditures in 2017 were 38.27% and 55.15% respectively out of total current health expenditures (5.05% of GDP ) in this region which is an insignificant amount (WDI, 2020). The average economic growth of this region in 2017 was 5.63% which improved over the year. Increased industrialization and urbanization may play a critical role in this improvement (WDI, 2020). In terms of energy utilization and carbon discharges the average per capita energy intake and CO\(_2\) emission in

\(^1\) The SAARC-BIMSTEC countries are Afghanistan, Bangladesh, Bhutan, India, Maldives, Nepal, Pakistan, Sri Lanka, Myanmar and Thailand.
2017 of this region were 697.61 kg of oil equivalent and 1.36 metric tons (WDI, 2020; BP, 2019; Knoema, 2020).

Some studies have been conducted (e.g. Jaba et al., 2014; Rahman et al., 2018; Duba et al., 2018; Hanif, 2018; Amuka et al., 2018; Shobande, 2020, among others) considering health status but they have proven to be inconclusive and are therefore of no help in formulating cohesive policies in the health sector. The differences of opinion may be due to the variety of methodologies, regions, countries, variables, and data, which create enigmatic nexus between related variables and health status, creating the opportunity for further investigations for formulating and executing proper policy strategies. In this regard our study is an effort to identify the decisive impact of health expenditure (both public and private), energy consumption and environmental pollution on health status by utilizing various voguish econometric tools in the case of the panel of SAARC-BIMSTEC countries.

The rationale for selecting the variables is that: (i) a combination of public and private health expenditures ensure better health status; (ii) energy consumption creates enough facilities to obtain good health status; (iii) excessive carbon emissions are detrimental to better health outcomes; (iv) economic growth offers improved amenities for safeguarding health status; and (v) effective sanitation facilities contribute to improved health status. More rationalization of the selection of variables is noted in section 3.1.

The key objectives of this study are:

i. To identify the impact of health expenditure (public and private), carbon emissions, energy consumption, sanitation facilities, and economic growth on health status in the SAARC-BIMSTEC region.

ii. To detect the long-run and short-run associations, causality and dynamics among health status, health expenditure (public and private), carbon emissions, energy consumption, sanitation facilities, and economic growth.

The major contributions of this study are: (i) this is the first study in the literature, to the best of our understanding, that investigates the effects of public and private health expenditures, energy consumption, carbon emissions, economic growth, and sanitation facilities on health status in the context of SAARC-BIMSTEC regions; (ii) this work
employs the latest available and inclusive data covering the period of 16 years (2002-2017); (iii) the results are attained by employing various erudite econometric tools e.g. cross-sectional dependence test, CADF panel unit root test, Pedroni and Kao test, Pool Mean Group (PMG) or panel autoregressive distributed lag (ARDL) method, Dumitrescu and Hurlin causality test; and (vi) the appraised outcomes will provide guiding principles for policy makers to ensure improved health status by espousing massive health expenditures (public and private), lower carbon emissions and energy consumption, and effective economic growth policies.

The study is designed in the following order: following the introduction, section 2 critically analyzes the literature review; section 3 explicates the data and methodology; section 4 provides and analyzes the estimated outcomes; and section 5 presents the conclusion and policy implications.

2. Literature Review

Contemporary literatures have made an effort to derive the most important factors relating to health and its nexus, which is not found irrefutable and ascertained to be relaxed in deferring from further investigations. Factors affecting health status are various. In this study, we will review the past literature under the following strands:

2.1 Health status - health expenditure nexus

Various researchers (Deshpande et al. (2014), Jaba et al. (2014), Duba et al. (2018), Bousmah et al. (2016), Becchetti et al. (2017), among others) noted the positive role of health expenditure on health indicators like life expectancy, infant and child mortality rate, and crude death rate. Deshpande et al. (2014) pointed out the significant positive effect of health care expenditure on life expectancy in developed countries but the insignificant impact in developing countries from the data of 181 countries. Similarly, Jaba et al. (2014) identified the positive impact of health expenditure on life expectancy in the case of 175 countries from the data of 1995-2010. In the same way, Duba et al. (2018) ascertained the positive effect of health care expenditure on life expectancy at birth for both male and female, considering 210 countries and regions over the period of 1995-2014. Bousmah et al. (2016) provided evidence that health expenditure positively affects health outcomes in the presence of high institutional quality in the context of 18 Middle East and North Africa (MENA) countries for the period of 1995–2012. Becchetti et al. (2017) also identified the
significant influence of health expenditure on reducing chronic diseases which lengthens life expectancy in the individual data of 13 countries.

In contrast, a number of researchers also found unclear or insignificant relationships between health expenditure and health indicators (see Rahman et al. (2018), Jiang et al. (2018), and Tran et al. (2017), among others). In this context Rahman et al. (2018) identified the positive impact of health expenditure on crude death rate and infant mortality rate but no significant impact on life expectancy at birth while conducting their study in the SAARC-ASEAN regions for the data period of 1995-2014. Tran et al. (2017) found that health care expenditure did not improve health and advocated reallocating public health expenditure from medical expenditures to social programs in the USA to raise life expectancy and reduce death. Jiang et al. (2018) found a varying effect of health care spending on life expectancy due to the heterogeneity of social development across the China over the period 2000-2010. Similarly, Hitiris and Posnett (1992), and Barlow and Vissandjee (1999) did not clearly observe the nexus between health expenditures and the length of life.

The ongoing debates also provide an unsettled idea about the emphasis should be on public or private health expenditure for ensuring better health facilities for people. In this context, Ray and Linden (2019) found a positive effect of public health expenditure but a negative effect of private health expenditure on the life expectancy at birth for 195 countries for the data period of 1995-2014 by employing GMM estimators. Rezapour et al. (2019) also revealed that there was a significant impact of public health expenditure but little or no significance of private health expenditure on life expectancy in the context of global countries. Similarly, Aísa et al. (2014) found a higher impact of public health expenditure but a lower impact of private health expenditure on life expectancy in the OECD countries over the period 1980-2000. Golinelli et al. (2017) revealed that higher public health expenditure lowered the infant mortality rate in Italy over the period 1999-2013. On the other hand, Nkemgha et al. (2020) identified the positive and statistically significant influence of private health expenditure but no significant impact of public health expenditure on life expectancy in Cameroon during 1980-2014. Similarly, van den Heuvel and Olariou (2017) declared that public health expenditure is not a key determinant of life expectancy while conducting a study in 31 European countries for the data period of 2014-2014. Many other researchers pointed out the positive role of both public and private health care expenditure on health status. For example, Novignon et al. (2012) presented the strong positive link of both public and private health care expenditure with life expectancy, had although public health spending had comparatively greater influence in case of 44 countries of sub-Saharan Africa from the panel data of 1995-2010. Akinci et al. (2014) also found that both the public and private health expenditures reduced maternal mortality rate, the under-five mortality rate, and the infant mortality rate in the MENA region during 1990 to 2010. Thus the inconclusive interconnection between health expenditure and health status demands further investigations.
2.2 Health status – energy consumption nexus

Energy is the life blood of the modern world ensuring better living for people by making technological amenities available, and providing sophisticated and improved medical facilities. Recently, some researchers (Shobande, 2020; Hanif 2018; Youssef et al., 2016; Jorgenson et al., 2014; among others) have tried to ascertain the link between health status and energy consumption. Shobande (2020) discovered that energy consumption reduces and carbon emission increases the child mortality in 23 African countries over the period of 1999-2014. Similarly, Hanif (2018) found that non-renewable energy consumption adversely affected life expectancy by raising the mortality rate, while renewable energy consumption reduced the mortality rate and extended life expectancy over the period of 1995-2015 in 34 Sub-Saharan African countries. In the same way, Youssef et al. (2016) found strong link between health and energy consumption in Africa, where in some African countries they found unilateral causality between energy use and life expectancy, and under-5 child mortality. Furthermore, Jorgenson et al. (2014) found that declining energy intensity and increasing energy efficiency improved human well-being in the 12 Central and Eastern European (CEE) countries for the period of 1992-2010. Thus, further investigation on the link between energy consumption and health status is required for correct policy initiatives.

2.3 Health status – environmental pollution nexus

The effect of environmental pollution on health is now becoming another key concern, and some research relates clearly to this vital issue (see Nathaniel and Khan, 2020; Duba et al., 2018; Amuka et al., 2018; Monsef and Mehrjardi, 2015; Mutizwa and Makochebanwa, 2015; and Sinha, 2014; among others. Nathaniel and Khan (2020) observed the negative but insignificant impact of carbon emissions on life expectancy in Nigeria for the study period 1970-2014. Duba et al. (2018) and Balan (2016) also found the negative impact of CO\textsubscript{2} emissions on life expectancy at birth. Mutizwa and Makochebanwa (2015) found that infant mortality increased due to environmental factors in 12 countries of Southern African Development Community (SADC) during the period 2000-2008.

On the other hand, Monsef and Mehrjardi (2015) showed that CO\textsubscript{2} emissions may not always be detrimental to health, specifically in developing countries because of lower level of emissions. Amuka et al. (2018) also found a positive linkage between carbon emissions and life expectancy in Nigeria over the period of 1995-2013. In addition, Sinha (2014) found a bidirectional causal association between CO\textsubscript{2} emissions and infant mortality rate in India which validated the feedback causality. Therefore, the study relating to the interconnection between environment and health status is alluring and deserves more thorough investigations in an erudite way.
From the exploration of the above mentioned literature, it has been observed that the results are unconvincing and not encouraging to the enunciation of any appropriate policy initiatives towards ensuring better health status for the people. Moreover, the combined consideration of both public and private health expenditures, energy consumption and environmental pollution on health status is absent in most of the literature, especially in the SAARC-BIMSTEC region. Therefore, this study will inevitably fill up the existing literary gaps and be pioneer for providing the expedient way of probing the impact of health expenditure, energy consumption and environmental pollution on health status and the articulation of proper policy implications.

3. Data and methodology

3.1. Theoretical or empirical rationale for choosing the variables

The theoretical justification for this study may be linked with various well-known theories e.g. human capital theory suggested by Gary S. Becker (1964) (Shobande, 2020; Breton, 2014; Na, 2012); the health care model recommended by Grossman (Shobande, 2020; Novignon et al., 2012; Grossman, 1972); and the inclusion of human capital in the neoclassical endogenous growth model by Romer (1996) (Novignon et al., 2012; Romer, 1996).

The empirical rationale for selecting our variables is mainly on the basis of the data availability, and past and existing literatures. We have utilized the variable life expectancy at birth to determine the health status following Lichtenberg (2014), Rahman et al. (2018), Ray and Linden (2019), Nkemgha et al. (2020), and others; public and private health expenditures are considered following Jaba et al. (2014), Duba et al. (2018), Rahman et al. (2018), and others; energy consumption in the line of Shobande (2020), Hanif (2018), Youssef et al. (2016), Jorgenson et al. (2014) among others; the carbon emissions to denote environmental pollution in the line of Nathaniel and Khan (2020), Duba et al. (2018), Amuka et al. (2018), Sinha (2014) among others; per capita GDP as a proxy of economic growth following Rahman et al. (2018), Shobande (2020), and others, and the sanitation facilities following Rahman et al. (2018) and Novignon et al. (2012) among others.
The identification of our used variables is as follows:

HS: Health status
PUH: Public health expenditure
PVH: Private health expenditure
ENC: Energy consumption
CO$_2$: Carbon emissions
PerGDP: Per capita gross domestic product
SAN: Sanitization facilities

The model used for empirical estimation is depicted below:

$$HS = f (PUH, PVH, ENC, CO_2, PerGDP, SAN)$$  \hspace{1cm} (1)

After converting all the variables of the equation $(1)$ into natural logarithmic form, we will obtain the elasticity of every variable from the coefficient values. Therefore, the equation $(1)$ can be depicted as:

$$LN_{HS_t} = \alpha + \beta_1 LN_{PUH_t} + \beta_2 LN_{PVH_t} + \beta_3 LN_{ENC_t} + \beta_4 LN_{CO_2_t} + \beta_5 LN_{PerGDP_t} + \beta_6 LN_{SAN_t} + \varepsilon_t$$  \hspace{1cm} (2)

Where, $\alpha$ is the intercept, and $\beta_1$, $\beta_2$, $\beta_3$, $\beta_4$, $\beta_5$, $\beta_6$ are coefficients and $\varepsilon_t$ is the error term.

### 3.2. Variables and data

To conduct our study we have utilized the health status as a dependent variable; as a proxy of this, variable life expectancy at birth is used. The independent variables are public and private health expenditures, carbon emissions, energy consumption, sanitization facilities, and economic growth. Both of the public and private health expenses are taken as a percentage of current health expenditures; energy consumption is considered to be a per capita oil equivalent, and the carbon emissions variable is taken as metric tons per capita. As the proxy variable for economic growth the per capita gross domestic product (GDP) is used, and the sanitation variable is considered as the usage of at least basic sanitation by the population ($\%$ of total population).

In this study we have used the data of 2002-2017 from the World Development Indicator (WDI, 2020) of the World Bank. But, due to the lack of data of energy use in case of
Afghanistan, Bhutan and the Maldives we have taken figures from the data (2002-2017) of Knoema’s (2020) data world atlas. For all other countries beyond WDI, carbon emission and energy consumption data for the period of 2015-2017 are collected from Knoema’s (2020) data world atlas and BP (2019) statistical review. The SAARC-BIMSTEC countries included in the study are Afghanistan, Bangladesh, Bhutan, India, Maldives, Nepal, Pakistan, Sri Lanka, Myanmar and Thailand.

3.3 Econometric approach
In this study we applied various econometric techniques for the purpose of estimation. We employed the following tests: a cross-sectional dependence test to observe the shock effect; Pesaran's (2007) CADF panel unit root test to confirm the stationarity of variables; the Pedroni (1999, 2004) and Kao (1999) cointegration tests to measure the equilibrium association; the panel pooled mean group (PMG) or panel ARDL method to estimate the long and short term equilibrium effects; and Dumitrescu and Hurlin’s (2012) heterogeneous panel causality to discover the direction of causality.

3.3.1 Cross-sectional dependence
Cross-sectional dependence of the variables may appear due to the prevalence of analogous geographic, economic, historical, ethnic and political jolts; hence these jolts should be checked before the detection of the unit root. In this study we have utilized four renowned cross-sectional dependency tests: Breusch and Pagan (1980) BP LM, Pesaran (2004) scaled LM, Pesaran (2004) CD, and Baltagi et al. (2012) biased-corrected scaled LM.

To examine the cross-sectional dependence among the panel data the below model is suggested by Breusch and Pagan (1980):

\[ CD_{BP} = \sum_{i=1}^{N-1} \sum_{j=i+1}^{N} \hat{p}_{ij}^{-2} \]

To overcome the limitations of the above, Pesaran (2004) develops the following LM statistics:
\[ CD_{LM} = \sqrt{\frac{1}{N(N-1)} \sum_{i=1}^{N-1} \sum_{j=i+1}^{N} (\hat{p}_{ij}^2 - 1)} \]

Pesaran (2004) recommends that if the cross-sectional size is greater than the time dimension, the following test statistic can be used instead.

\[ CD = \sqrt{\frac{2T}{N(N-1)} \sum_{i=1}^{N-1} \sum_{j=i+1}^{N} \hat{p}_{ij}^2} \]

From the simple asymptotic bias correction, the scaled LM test as recommended by Baltagi et al. (2012) is:

\[ CD_{BC} = \sqrt{\frac{1}{N(N-1)} \sum_{i=1}^{N-1} \sum_{j=i+1}^{N} (\hat{p}_{ij}^2 - 1) - \frac{N}{2(T-1)}} \]

Where \( \hat{p}_{ij} \) specifies a correlation among the errors. In this test the null hypothesis is \( H_0 \): no cross-sectional dependence whereas the alternative hypothesis is \( H_1 \): cross-sectional dependence.

### 3.3.2 Panel unit root test

Detection of unit root in the series of variables is essential to estimate it in the PMG or panel ARDL model. In this regard we have used the cross-sectional augmented Dickey-Fuller (CADF) test because of the consideration of cross-sectional panel data following the methodology of Pesaran (2007) under the following equation:

\[ \Delta x_{it} = \alpha_i + p_1 x_{it-1} + \delta_1 \bar{x}_{t-1} + \sum_{j=0}^{n} \eta_{ij} \Delta \bar{x}_{it-1} + \sum_{j=0}^{k} \Psi_{ij} \Delta x_{it-1} + \varepsilon_{it} \]

Where, \( \bar{x}_{t-1} \) and \( \Delta \bar{x}_{it-1} \) indicate the cross-sectional averages of lagged levels and first difference individual series respectively.
3.3.3 Co-integration test

For co-integration among the variables, a residual based cointegration test as advised by Pedroni (1999, 2004) has been employed on the basis of two varieties: panel test and group test. The panel test has four statistics under the within dimension: panel PP, panel rho, panel \(v\) and panel ADF, whereas the group test has three statistics under the between dimension: group PP, group rho and group ADF. It is considered that the mentioned statistics are standard, normally and asymptotically distributed residuals based on the following long-run model:

\[ Y_{it} = a_t + \zeta_t + \sum_{j=1}^{m} \Psi_{jt} X_{jit} + \varepsilon_{it} \]

The construction of the appraised residuals is as under:

\[ \varepsilon_{it} = p_i \varepsilon_{it-1} + u_t \]

The comparison is made based on the maximum likelihood panel cointegration statistics considering no co-integration among the variables as the null hypothesis. The cointegration system for panel data recommended by Pedroni (1999, 2004) is as follows:

\[ Y_{it} = a_i + \Psi X_{it} + \varepsilon_{it} \]

We have also used another co-integration test to observe homogenous nexus as suggested by Kao (1999) under two types of test: the Dickey-Fuller and the Augmented Dickey Fuller, where no co-integration among the variables as the null hypothesis is assumed.

3.3.4 Pool mean group or panel ARDL estimation

To identify the dynamic nexus between health status and public and private health expenditures, carbon emissions, energy consumption, sanitation facilities, and economic growth we have utilized the pool mean group (PMG) or panel autoregressive distributed lag (ARDL) method under the following equation:

\[ \Delta(Y_{it}) = \sum_{j=1}^{p-1} \eta^i_j \Delta(Y_{jt-1}) + \sum_{j=1}^{q-1} \rho^i_j \Delta(X_{jt-1}) + \eta^i [ (Y_{jt-1} - \psi^i_0 + \psi^i_0 (X_{jt-1})) ] + \varepsilon_{it}, \]
Where, Y is the health status. X is the set of explanatory variables comprising public and private health expenditures, energy use, CO₂ emissions, economic growth, and sanitation facilities. The short-term coefficients of dependent and independent variables are symbolized respectively by η and ρ, where the long-term coefficients is denoted by Ψ. The speed of adjustment towards long-term affiliation is denoted by η0, and the time-varying error term by ε. We have used the subscript i to determine country and t for time. The long-term growth regressions are shown within the square brackets.

3.3.5 Heterogeneous panel causality test
To perceive the short-term bivariate causal association among the variables we have adopted the Dumitrescu and Hurlin (2012) heterogeneous panel causality test that considers the cross-sectional dissimilarities under the following model:

\[ y_{it} = a_t + \sum_{j=1}^{J} \eta_{i}^{(j)} y_{i,t-j} + \sum_{j=1}^{J} \rho_{i}^{(j)} x_{i,t-j} + \varepsilon_{i,t} \]

The \( x_{i,t} \) and \( y_{i,t} \) denote the observations of two stationary variables of individual i in period t, \( j \) exposes the lag length, \( \eta_{i}^{(j)} \) displays the autoregressive parameter, and \( \rho_{i}^{(j)} \) designates the regression coefficient varying within the groups. The identical lag order J for all individuals is considered so that the panel may be balanced. It is a fixed coefficient, normally distributed and permits the case of heterogeneity.

This test assumes no causality as the null hypothesis and the existence of causality between variables as the alternative hypothesis, as under:

\( H_0 : K_i = 0 \quad \forall i=1, \ldots, N \)

\( H_1 : K_i = 0 \quad \forall i=1, \ldots, N_1 \)

\( K_i \pm 0 \quad \forall i=N_1 + 1, N_1 + 2, \ldots, N \)

Here, \( N_1 \) represents the unknown parameter that fulfils the condition \( 0 \leq N_1 / N < 1 \). The ratio of \( N_1/N \) ought to be lower than 1, because \( N_1 = 0 \) indicates causality for all individuals in the panel, whereas, \( N_1 = N \) means no causality, so that we failed to reject the null hypothesis.
4. Results and discussions

4.1 Descriptive statistics

In Table-1, the results of the descriptive statistics of our considered variables are displayed in natural logarithmic form. We have found that the mean, median, maximum, minimum, and standard deviation values of the health status are 4.224, 4.218, 4.361, 4.039, and 0.076 respectively. The mean values of public health expenditure, private health expenditure, energy consumption, CO\textsubscript{2} emissions, per capita GDP, and sanitation facilities are, consecutively, 3.263, 3.966, 6.080, -0.398, 7.289, and 4.016, and their standard deviations are, consecutively, 0.827, 0.494, 1.032, 1.093, 0.893, and 0.431.

Table 1: Descriptive statistics

| Variables | LNHS  | LNPUH | LNPVH | LNENC | LNCO2 | LNPERGDP | LNSAN |
|-----------|-------|-------|-------|-------|-------|----------|-------|
| Mean      | 4.224 | 3.263 | 3.966 | 6.080 | -0.398| 7.289    | 4.016 |
| Median    | 4.218 | 3.231 | 4.221 | 6.120 | -0.406| 7.060    | 4.055 |
| Maximum   | 4.361 | 4.336 | 4.455 | 7.700 | 1.530 | 8.963    | 4.599 |
| Minimum   | 4.039 | -0.115| 2.661 | 2.979 | -3.259| 5.800    | 2.989 |
| Std. Dev. | 0.076 | 0.827 | 0.494 | 1.032 | 1.093 | 0.893    | 0.431 |
| Observations | 160 | 160 | 160 | 160 | 160 | 160 | 160 |

Note: All the variables are converted into the natural logarithm form.

4.2 Cross-sectional dependence test results

We have reported the values of four types of cross-sectional dependence tests and their probabilities in Table 2. All the variables except public health expenditure have confirmed the significance in all four tests, which has assured three tests. From the Table-2 we can observe that the probability values of our studied variables are at different levels and can therefore reject the null hypothesis of cross-sectional independence.

Table 2: Cross-sectional dependence test results

| Variables | Breusch-Pagan LM | Pesaran scaled LM | Bias-corrected | Pesaran CD |
|-----------|------------------|-------------------|---------------|------------|
|           |                  |                   |               |            |
4.3 Unit root test results

To detect the presence of the unit root we have employed the cross-sectional ADF unit root test (CADF) as per the methodology of Pesaran (2007), due to the contemplation of the cross-sectional dependence issue. By applying the constant only we have found that the LNHS, LNPUH, and LNSAN have no unit root at level, where the LNENC, LNCO₂, LNPVH, LNPERGDP have unit root at level but no unit root at their first difference (Table-3). As per the methodology of Pesaran et al. (2001), we may utilize the ARDL model in case the series has no unit root or stationary at level I(0), at first difference I(1) or both, but none may be stationary at secondary difference I(2).

Table 3: Unit root test results

| Variables   | Constant | Order of Integration |
|-------------|----------|---------------------|
|             | Level    | p-value | 1st difference | p-value |           |
| LNHS        | -5.248***| 0.000   | -2.229*        | 0.066   | I(0)     |
| LNENC       | -1.840   | 0.368   | -3.272***      | 0.000   | I(1)     |
| LNCO₂       | -1.578   | 0.680   | -3.381***      | 0.000   | I(1)     |
| LNPVH       | -2.399** | 0.020   | -3.772***      | 0.000   | I(0)     |
| LNPERGDP    | -2.047   | 0.165   | -3.313***      | 0.000   | I(1)     |
| LNSAN       | -3.189***| 0.000   | -2.335**       | 0.035   | I(0)     |

Note: ***, **, and * denote significance level at 1%, 5% and 10%, respectively.
4.4 Cointegration test results

To investigate the prevalence of the long-run nexus between health status and health expenditure (both public and private), energy consumption, environmental pollution, economic growth, and sanitation we have taken the assistance of two kinds of panel cointegration tests as proposed by Pedroni (1999, 2004) and Kao (1999). Both tests have assured the existence of long-term cointegration among our studied variables, and are therefore suitable to evaluate the parameters of the dynamic error correction model by applying the panel ARDL model or the PMG method (Table-4).

Table 4: Pedroni and Kao panel cointegration test results

| Pedroni cointegration test | Statistic | Prob. | Weighted Statistics | Prob. |
|----------------------------|-----------|-------|---------------------|-------|
| *Alternative hypothesis: common AR coefficients (within-dimension)* |           |       |                     |       |
| Panel v-Statistic          | -0.837    | 0.799 | 0.347               | 0.3642|
| Panel rho-Statistic        | 3.456     | 0.999 | 2.644               | 0.996 |
| Panel PP-Statistic         | 0.522     | 0.699 | -1.614*             | 0.053 |
| Panel ADF-Statistic        | -1.821**  | 0.034 | -2.554***           | 0.005 |

*Alternative hypothesis: individual AR coefficients (between-dimension)*

| Group rho-Statistic        | 3.843     |       | 0.999               |       |
| Group PP-Statistic         | -3.082*** |       | 0.001              |       |
| Group ADF-Statistic        | -2.995*** |       | 0.001              |       |

Kao cointegration test

| Statistic     | t-statistic | Prob. |
|---------------|-------------|-------|
| ADF           | -3.593***   | 0.000 |

Note: ***, ** and * denote significance level at 1%, 5% and 10%, respectively.

4.5 The results of Panel ARDL estimation

From the model of PMG or panel ARDL(1,1,1,1,1,1,1) and applying Akaike information criterion in the case of constant (level) case-II we found the long-run and short-run association among our studied variables (Table-5). We observed that the coefficients of the public and private health expenditures and energy consumption are, respectively, 0.014, 0.030, and 0.027, which are positive and statistically significant at 1% level. The coefficient
of environmental pollution proxied by CO2 is -0.085 which is negative and statistically significant at 1% level. The coefficients of economic growth and sanitation facilities are 0.029 and 0.018, where the former is statistically significant at 5% level but the latter is insignificant. The coefficient of error correction term (COINTEQ01) is -0.034, which is negative and statistically significant even at 1% level assuring the long-run cointegration among the considered variables. However, in the short-run we found that neither variable is statistically significant, which denotes that the fruit of the above variables on health status cannot be found instantaneously and will have to wait for the long run. The obtained outcomes are pragmatic and rational, which has remarkable empirical significance in the health sector.

Table 5: PMG or panel ARDL results

| Variable   | Coefficient | Std. Error | T-Statistic | Prob.  |
|------------|-------------|------------|-------------|--------|
| Long Run   |             |            |             |        |
| LNPUH      | 0.014***    | 0.005      | 2.691       | 0.009  |
| LNPVH      | 0.030***    | 0.011      | 2.821       | 0.006  |
| LNENC      | 0.027***    | 0.009      | 3.095       | 0.003  |
| LNCO2      | -0.085***   | 0.024      | -3.568      | 0.001  |
| LNPERGDP   | 0.029**     | 0.012      | 2.368       | 0.021  |
| LNSAN      | 0.018       | 0.069      | 0.261       | 0.795  |
| Short Run  |             |            |             |        |
| COINTEQ01  | -0.034***   | 0.012      | -2.848      | 0.0057 |
| D(LNPUH)   | -0.001      | 0.001      | -0.702      | 0.485  |
| D(LNPVH)   | -0.005      | 0.004      | -1.240      | 0.219  |
| D(LNENC)   | 0.0004      | 0.0004     | 1.021       | 0.311  |
| D(LNCO2)   | 0.003       | 0.002      | 1.616       | 0.110  |
| D(LNPERGDP)| -0.007      | 0.004      | -1.604      | 0.113  |
| D(LNSAN)   | -1.122      | 1.019      | -1.101      | 0.274  |
| Constant   | 0.124**     | 0.048      | 2.587       | 0.0117 |

Note: ***,**, and * denote significance level at 1%, 5% and 10%, respectively.

4.6 Country specific short-term test results

The outcomes of the country specific tests in the short-run are displayed in Table-6. The negative effect of health care expenditure in various countries may be due to the inefficiency of spending in the health sectors and rampant corruption in health financing. In the short-run,
the positive consequence of CO₂ emissions on the health status of people in some countries may be due to the externality effect of economic development which may generate huge income at the cost of the environment and people may be able to afford better food and housing, but in the long-run they may affect people’s health negatively. For some countries the negative coefficients of energy consumption, economic growth and sanitation facilities in the short-run may be due to the immediate adverse impact of these variables by creating pressure on the cost, environment and inefficient utilization of energy and sanitation.

Table 6: Short-term country specific results of all sample countries

| Country     | COINTEQ01 | LNPUH       | LNPVH       | LNENC       | LNCO2       | LNPERGDP    | LNSAN       | Constant    |
|-------------|-----------|-------------|-------------|-------------|-------------|-------------|-------------|-------------|
| Afghanistan | 0.012***  | -0.001***   | 0.001***    | 0.001***    | 0.001***    | 0.523***    | -0.055***   |             |
| Bangladesh  | -0.03***  | -0.003***   | -0.016***   | 0.001***    | 0.006***    | 0.018***    | 0.035**     | 0.098***    |
| Bhutan      | 0.0121*** | 0.005***    | -0.0002***  | 0.003***    | -0.0001***  | -0.003***   | 0.877***    | -0.052***   |
| India       | -0.080*** | 0.003***    | 0.009***    | -0.001***   | 0.002***    | -0.010***   | -0.089***   | 0.302***    |
| Maldives    | -0.058*** | 0.001***    | -0.001***   | -0.001***   | 0.002***    | -0.001***   | 0.017***    | 0.218***    |
| Nepal       | -0.006*** | -0.0002***  | -0.001***   | -0.000***   | 0.0002***   | -0.001***   | 0.017***    | 0.024***    |
| Pakistan    | -0.066*** | 0.00003***  | -0.002***   | 0.00003***  | -0.002***   | -0.014***   | -0.098***   | 0.241***    |
| Sri Lanka   | -0.090*** | -0.004***   | -0.002***   | 0.0002***   | 0.017***    | -0.036***   | -6.245      | 0.378***    |
| Myanmar     | -0.020*** | 0.0003***   | 0.00002***  | -0.0004***  | 0.001***    | -0.018***   | -7.951*     | 0.032***    |
| Thailand    | -0.014*** | -0.010***   | -0.006***   | -0.001***   | 0.002***    | -0.008***   | 1.689**     | 0.021***    |

Note: *** and ** denote significance level at 1%, 5% and 10%, respectively.

4.7 Results of the panel causality test of Dumitrescu and Hurlin (2012)

In Table-7 the Dumitrescu and Hurlin (2012) panel causality outcomes are shown by W-statistic and corresponding probability. In this study, we have observed the bidirectional causality between health status, and health expenditures (both public and private), carbon emissions, economic growth, and sanitation facilities, but unidirectional causality emanating from the energy use to health status.

Table 7: Causality test results

| Null Hypothesis                  | W-Stat. | Prob. | Decision                      |
|----------------------------------|---------|-------|--------------------------------|
| LNPUH does not cause LNHS        | 3.495***| 0.000 | LNPUH ↔ LNHS (bidirectional causality) |
| LNHS does not cause LNPUH        | 2.783** | 0.012 |                                |
| LNPVH does not cause LNHS        | 3.786***| 0.000 | LNPVH ↔ LNHS (bidirectional causality) |
| LNHS does not cause LNPVH        | 2.454** | 0.046 |                                |
| LNENC does not cause LNHS        | 5.315***| 0.000 | LNENC → LNHS (unidirectional causality) |
| LNHS does not cause LNENC        | 1.623   | 0.501 |                                |
|                                | Value  | p-value | Notes                                      |
|--------------------------------|--------|---------|--------------------------------------------|
| LNCO2 does not cause LNHS      | 9.660  | 0.000   | LNCO2 ↔ LNHS (bidirectional causality)     |
| LNHS does not cause LNCO2      | 3.265  | 0.000   |                                             |
| LNPERGDP does not cause LNHS   | 9.837  | 0.000   | LNPERGDP ↔ LNHS (bidirectional causality)  |
| LNHS does not cause LNPERGDP   | 3.094  | 0.003   |                                             |
| LNSAN does not cause LNHS      | 19.951 | 0.000   | LNSAN ↔ LNHS (bidirectional causality)     |
| LNHS does not cause LNSAN      | 37.283 | 0.000   |                                             |

Note: '***', '**', and '*' denote significance level at 1%, 5% and 10%, respectively.

5. Conclusion and policy implications

This paper explored the nexus between health status and health expenditures (both public and private), energy consumption and environmental pollution in the SAARC-BIMSTEC region. Using annual data from 16 years (2002-2017), we have utilized the panel autoregressive distributed lag (ARDL) model, heterogeneous panel causality test, cross sectional dependence test, cointegration test and Pesaran cross sectional dependent (CADF) unit root test to estimate the findings. The acquired results authorize the cointegration among the variables used; public and private health expenditure, energy consumption, and economic growth have positive and statistically significant influence, and environmental pollution has negative and significant effect on the health status of the region in the long-run, but no panel wise significant impact is revealed in the short-run. A two-way causal nexus between health status and environmental pollution, public and private health expenditure, economic growth and sanitation facilities, and a one-way causality running from energy use to health status have also been found. The attained results are theoretically and empirically consistent, and have significant policy implications in the health sector. The policy implication of the findings is: the better and improved health status required to be safeguarded by enunciating the effective policies on both public and private health expenditures, environmental pollution, energy consumption, and economic growth. In this context the following recommendations should be stressed:

i. **Public-private health expenditure policy mix:** In order to ensure better health care facilities, a relaxed and affordable health expenditure policy combining public and private initiatives is needed. Establishing a large number of hospitals, supplying a good number of efficient doctors, nurses, medical stuffs and arranging modern health
care facilities and better diagnosis instruments—all are essential to guarantee better health. All will be sustainable by the use of an adequate, corruption free and erudite public-private health expenditure policy mix.

ii. *Efficient energy policy:* Energy provides the impetus of efficient and ostentatious economic growth which also affects the health status of the people through different channels. By improved living standards, supplying adequate electricity, facilitating manufacturing new drugs, delivering different health related facilities through different ways, energy plays a crucial role in the modern health sector. In this regard the negative externality of energy production is also a matter of concern because different diseases are spread through environmental pollution. Therefore, an efficient, affordable, reliable, modern energy policy is urgently required by emphasizing the use of clean and renewable energy in the region.

iii. *Reducing environmental pollution:* Environmental pollution affects human health adversely by generating different viruses and bacteria, which are responsible for various diseases like bronchitis, heart problems, lung problems, flu related diseases like COVID-19 and many other potentially fatal conditions. In this regard a green environment ensuring by the reduction of CO₂ emissions should be encouraged. Proper policy regarding the clean energy, green development, and reforestation activities may curb environmental pollution and foster longevity.

iv. *Health friendly economic growth policy reforms:* Economic growth and development enable people to earn more and benefit from more modern health facilities; however, the negative external effect of growth activities (increased pollution) may be detrimental to public health. In this context sustainable, conducive, and smart health friendly economic growth policy reforms will be helpful to increase people’s life spans throughout the world, and more particularly in the countries that have been mentioned above.
Declarations:

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Availability of supporting data
Data used in this study are collected from the World Development Indicator (WDI, 2020) of the World Bank, Knoema’s (2020) data world atlas, and BP (2019) statistical review. The complete structured data set used for the study may be supplied by the corresponding author on reasonable request.

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MMR contributed in conceptual and methodological development, variable selection, result analysis, writing abstract, polishing and editing, and improving the quality of the manuscript, and overall carefully supervising. KA made study plan, literature review, data collection, writing main sections of the paper, econometric estimation, and data and result analysis, undertaking the responsibility of corresponding author of this paper.

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