The Gap on GAP: Tackling the Problem of Differing Data Distributions in Bias-Measuring Datasets

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Abstract

Diagnostic datasets that can detect biased models are an important prerequisite for bias reduction within natural language processing. However, undesired patterns in the collected data can make such tests incorrect. For example, if the feminine subset of a gender-bias-measuring coreference resolution dataset contains sentences with a longer average distance between the pronoun and the correct candidate, an RNN-based model may perform worse on this subset due to long-term dependencies. In this work, we introduce a theoretically grounded method for weighting test samples to cope with such patterns in the test data. We demonstrate the method on the GAP dataset for coreference resolution. We annotate GAP with spans of all personal names and show that examples in the female subset contain more personal names and a longer distance between pronouns and their referents, potentially affecting the bias score in an undesired way. Using our weighting method, we find the set of weights on the test instances that should be used for coping with these correlations, and we re-evaluate 16 recently released coreference models.

1 Introduction

AI systems trained on biased or imbalanced data can propagate and amplify observed patterns and make biased decisions at the time of evaluation and deployment Mehrabi et al. [2019]. To detect the underlying bias in released natural language processing (NLP) systems and to increase their fairness, several diagnostic datasets have been introduced, commonly focusing on gender bias Rudinger et al. [2018], Zhao et al. [2018], Webster et al. [2018], Kiritchenko and Mohammad [2018]. In addition to overall performance, these works also define the bias score of the evaluated model, usually the difference or ratio between the performance of the model on the groups of interest, e.g., female and male subsets of the data. However, when the subsets corresponding to the groups of interest consist of data coming from different distributions, undesired imbalances may appear, leading to inaccurate bias scores. For example, Webster et al. [2018] constructed the GAP coreference dataset by collecting examples from English Wikipedia and observe that a random baseline does not achieve a balanced bias score, despite being unbiased by design. They note that feminine sentences in the dataset contain more personal names and therefore more distractor mentions. If a coreference model is negatively affected by a larger number of potential candidates, it could appear more biased against female examples than it actually is. In this work, we introduce a method for coping with imbalances in such bias-detection datasets. We tailor the demonstration around gender bias, however, the method can be applied to any other type of bias as well.

A possible solution to the problem of imbalances caused by different data distributions could be the augmentation of the data by introducing examples with swapped genders. While this method has

1The annotations, the computed weights on the GAP test set, and the code can be found at https://github.com/vid-koci/weightingGAP

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been applied to training data [Zhao et al., 2018], for test data, it could have an unforeseen impact on different NLP systems. For example, such data augmentation may introduce instances that contain biologically or historically inaccurate facts, such as “men giving birth” or “historical figures being of the opposite gender”. Since GAP was collected from Wikipedia, a large number of examples with swapped gender could suffer from this problem.

Alternatively, examples can be constructed from manually crafted templates where genders can be swapped without risking to introduce inaccurate facts [Rudinger et al., 2018, Zhao et al., 2018, Kiritchenko and Mohammad, 2018]. While such tests of bias are important due to their controlled environment, they are synthetic and hence may not give a full picture of the underlying biases that can emerge when the model is used in practice.

The method introduced in this work assigns weights to test samples to cope with undesired imbalances in bias-measuring datasets. Given a list of properties that should not correlate with the gender of data examples, we derive a set of linear equations that should hold for the weighting of the test samples. At the same time, we minimize the likelihood of introducing additional noise into the measured bias of the models, by deriving an optimization objective that minimizes the upper bound of such noise. The search for optimal weights under given constraints is then implemented as a linear program.

We demonstrate the use of this approach on the GAP dataset for coreference resolution [Webster et al., 2018], the currently largest dataset for gender-bias-measuring. We annotate the GAP test set with all mentions of personal names. The annotations will be publicly released and may potentially be used for other research directions, such as the evaluation of named-entity recognition (NER) systems.

We show that, in the GAP test set, the feminine examples contain, on average, 0.75 more names per sentence than the masculine ones. Similarly, the correct candidate usually stands 0.5 candidates further away from the pronoun in feminine examples than in masculine ones. These are all imbalances that can affect the score of a model. We show the effectiveness of our weighting method by showing that a series of unbiased baselines indeed achieve scores closer to the unbiased score on the weighted test set. Finally, we re-evaluate 16 recently released coreference models on the weighted test set of GAP. We observe that several models change their bias scores when evaluated on the weighted test set, although most of these changes are small. We encourage future research to use the introduced weighted bias metric instead.

## 2 Weighting Method

In this section, we present our weighting method. While, for ease of presentation, we describe our method on gender-bias detection, the method can be applied to any type of bias detection, and it also generalizes to tasks where one needs to detect biases among $n > 2$ classes, e.g., racial bias, by observing every pair of classes separately. The current version of the method assumes that accuracy is used as a metric of performance. We leave the analysis of other potential metrics to future work.

### 2.1 Definitions and Objectives

Let $D$ be a bias-testing dataset with $n$ examples $D = \{x_1, \ldots, x_n\}$. Let $A$ and $B$ be non-overlapping subsets of $D$. We assume that $A \cup B = D$, i.e., we ignore examples outside of observed sets, if any. The aim is to compare the performance of a model on $A$ and $B$, in order to see if the model is biased.

Let $S_1, \ldots, S_m$ be subsets of $D$, such that $S_j$ consists of all examples with a property that is not specific to the sets $A$ and $B$ but that could have an impact on the performance of an evaluated model. For example, in the context of coreference resolution, one of the observed properties can be the number of referents in an example. In such an example, one set $S_k$ would consist of all examples with exactly $k$ potential referents. Note that these sets may overlap, as properties do not have to be mutually exclusive. We assume that these properties and sets are explicitly identified beforehand, and we refer to them as identified properties.

Let $C \subseteq D$ be a set of examples that a model solves correctly. Generally, the accuracy of a model corresponds to $|C| / |D|$. However, the performance of a model and hence $C$ may have a significantly different overlap with $S_j$ than with $D \setminus S_j$. Less formally, a model may be more/less likely to solve examples in the set $S_j$. To obtain an accurate bias measure, properties that do not influence the bias
should be evenly distributed across $A$ and $B$. If this is not the case for a bias-detection dataset, we adapt the bias metric so that $S_i \cap A$ carries equal weight as $S_i \cap B$ in the final score.

To achieve this, we assign to each example $x_i$ its weight $w_i$ and replace the accuracy with the weighted accuracy. We aim to find a set of weights $W$, such that $\sum_{x_i \in A \cap S_j} w_i = \sum_{x_i \in B \cap S_j} w_i$. Additionally, we impose the following restrictions on the weights:

- Balance between the observed sets: $\sum_{x_i \in A} w_i = \sum_{x_i \in B} w_i$.
- Fixed sum: $\sum_{i=1}^n w_i = n$.
- Non-negativity: $w_i \geq 0$, for all $i \in \{1, \ldots, n\}$.

The first two will simplify the future derivations, while the last one is put in place to avoid the situation where an incorrect answer is preferred over the correct one. A direct consequence of the first two is that the sum of all weights of one gender is fixed to $\sum_{x_i \in A} w_i = \sum_{x_i \in B} w_i = \frac{n}{2}$.

There could exist several sets of weights that meet the above criteria. Among them, we prefer the distribution that minimizes the potential exacerbation of other patterns in the data, that is, any changes in the bias score of a model that are not directly related to the above-identified properties. Let $\text{Acc}_W^A(C)$ and $\text{Acc}_W^B(C)$ be the unweighted and weighted accuracy, respectively, obtained by a set of correct answers $C$ on a set $D$. Since bias scores compare the performance on both $A$ and $B$, we aim to minimize both

\[
(1) \ |\text{Acc}_A^A(C \cap A) - \text{Acc}_A^A(C \cap A)| \quad \text{and} \quad (2) \ |\text{Acc}_W^B(C \cap B) - \text{Acc}_W^B(C \cap B)|
\]

for any $C \subseteq D$. This objective covers two cases:

- When $C$ corresponds to correct answers of a model, we minimize the difference in weighted and unweighted accuracy on the sets $A$ and $B$.
- When $C$ is a set of examples with some property other than the ones captured by $S_1, \ldots, S_m$, we aim to retain its original overlap with $A$ and $B$. The overlap between sets with unidentified properties and the sets $A$ and $B$ should not be removed, as they may be an important indicator of the underlying bias. An example of such a property in the context of gender bias in NLP is the amount of out-of-vocabulary words, which could be larger for feminine examples, should the text about men be more prevalent in the data used to construct the vocabulary.

Our method minimizes the upper bound on the differences (1) and (2). By considering the upper bound rather than the average case, we avoid making assumptions about the distribution of $C$.

**Theorem 2.1 (Upper Bound on Introduced Noise).** To minimize the upper bounds of

\[
|\text{Acc}_W^A(C \cap A) - \text{Acc}_A^A(C \cap A)| \quad \text{and} \quad |\text{Acc}_W^B(C \cap B) - \text{Acc}_W^B(C \cap B)|
\]

for any unknown set $C \subseteq D$, it is sufficient to minimize

\[
\sum_{x_i, x_j \in A} \max(w_i, w_j) + \sum_{x_i, x_j \in B} \max(w_i, w_j).
\]

The full proof is given in Appendix A.

### 2.2 Solving the Optimization Problem

All listed conditions and criteria can be phrased as a linear program. Balance between the subsets, fixed sum, non-negativity, and removing correlations are linear constraints. The optimization objective can be phrased as a linear function by introducing auxiliary variables $m_{i,j}; 1 \leq i, j \leq n$ for $\max(w_i, w_j)$. The following constraints have to hold for each of them: $m_{i,j} \geq w_i$ and $m_{i,j} \geq w_j$.

To summarize, we collect all derived constraints for the linear program:

- $\sum_{x_i \in A} w_i = \sum_{x_i \in B} w_i$.
- $\sum_{i=1}^n w_i = n$. 

\[ w_i \geq 0 \text{ for all } i \in \{1, \ldots, n\}. \]

\[ \sum_{x_i \in A \cap S_j} w_i = \sum_{x_i \in B \cap S_j} w_i \text{ for all } S_j. \]

For all \( i, j \) such that \( i < j \) and either \( w_i, w_j \in A \) or \( w_i, w_j \in B \): \( m_{i,j} \geq w_i \) and \( m_{i,j} \geq w_j \).

The criterion function is equal to
\[
\min \sum_{x_i, x_j \in A \atop i > j} m_{i,j} + \sum_{x_i, x_j \in B \atop i > j} m_{i,j}.
\]

A linear-program solver can then be used to find the minimum to this function.

3 Experiments

In this section, we demonstrate the use of our weighting method on the GAP dataset [Webster et al. 2018]. First, we show that feminine examples contain more candidates than masculine examples, and that the correct candidate usually stands further away from the pronoun in feminine examples than in masculine ones. We show that weighting solves these imbalances, as several unbiased baselines obtain scores closer to 1 after weighting (1 is a balanced score). Finally, we re-evaluate 16 publicly released models for coreference resolution, observing that the majority of these models were only slightly affected by these properties.

3.1 The GAP Dataset

GAP is a corpus of challenging examples of pronouns from English Wikipedia. It was introduced as a gender-balanced dataset, so that exactly half of the pronouns are masculine, and half are feminine [Webster et al. 2018]. The test set, which the rest of the paper is about, consists of 2000 text spans. The dataset comes with a development and validation set; however, they are not the focus of this work. For each text span, one pronoun has to be resolved. Pronoun resolution is treated as a binary classification task, with the goal to determine whether a single candidate is the referent of the pronoun or not. Note that candidates are not given as input and the model is expected to find them on its own. It is guaranteed that candidates are always personal names from the input text and that at most one of them is the correct referent. [Webster et al. 2018] define a bias measure as ratio between the \( F_1 \)-scores on the feminine and masculine subsets, \( F_1^F / F_1^M \). An unbiased system is therefore expected to achieve a bias score around 1. An example from GAP can be found below:

**Kathleen first appears when Theresa and Myra visit her in a prison.**

*Kathleen: True, Theresa: False*

During the scoring, the output of any evaluated model is compared to two candidates, specified by the example.

Note that any incorrect candidate adds noise to the bias score. If a model answers *Theresa*, it will be penalized with an additional false-positive outcome, unlike a model that answered *Myra*, despite both being equally wrong. Since there is never more than one correct candidate per sentence, and the candidates are not known in advance, comparing the prediction only with the correct candidate is thus not just sufficient, but also a more accurate bias measure. Thus, for measuring bias, we replace the \( F_1 \) score with accuracy, commonly used as a performance metric in coreference resolution [Emami et al. 2019, Rahman and Ng 2012, Sakaguchi et al. 2019].

To be able to observe the effect of our weighting method, we first introduce a plain accuracy-based bias metric **acc-Bias**. We measure the accuracy on positive candidates in the masculine subset \( A_M \) and the accuracy on positive candidates in the feminine subset \( A_F \) and define acc-Bias as \( A_F / A_M \). Results of this metric will be compared to a later-introduced weighted accuracy. Text spans with no positive examples are dropped, reducing the size of the test set by approximately 10%.

A possible improvement of GAP that we do not address in this work is a more fine-grained analysis and stricter definition from a linguistic perspective. The motivation by [Webster et al. 2018] is focused on biosocial gender, that is, comparing performance of models when the candidates are masculine and when they are feminine. However, in practice, the dataset measures the impact of grammatical gender, as the author define the gender of an example to match the gender of the pronoun in question. While these two types of gender largely overlap in English, mismatch can happen, e.g., in the case of...
personalization or misgendering. We refer to [Ackerman, 2019] for a detailed definition, comparison of these types of gender, and the analysis of the mismatches.

We found examples containing such mismatch not to have a strong presence in the GAP test set, however, the exact number is hard to estimate due to the lack of context in many of the examples. We highlight that addressing this is necessary before the dataset-creating approach by [Webster et al., 2018] can be used on languages with a stronger presence of grammatical gender, e.g., Russian and German.

3.2 Baselines

We re-implement the random and token distance baselines introduced by [Webster et al., 2018]. First, we find all personal names in the input text using an off-the-shelf named entity recognition (NER). Each baseline is implemented with two NER systems: Google Cloud NL API and Spacy en_core_web_lg, abbreviated Spacy-lg. Additionally, as we have manually labeled all spans in the GAP test set that correspond to a personal name, we use these annotations to implement Ground-Truth baselines, which are thus not affected by potential mistakes of the NER systems.

In the random baseline implementation, a random personal name is picked from the list. Note that our implementation of the random baselines exhibits a different performance than the one from Webster et al. [2018]. They report adding heuristics to eliminate obviously incorrect candidates; we do not follow them to avoid adding any noise.

In the token distance baseline implementation, the personal name closest to the pronoun is selected. Distance is measured in number of tokens, using the Spacy tokenizer. We rename this baseline as Dist-1 baseline and introduce Dist-2 and Dist-3 baselines, where we pick the second closest and third closest personal name, respectively. If there are fewer than 2 or 3 candidates in the sentence, then we consider all answers to be False, that is, we give no answer. We do not introduce higher-order distance-based baselines. Their accuracy drops and with it the denominator in the bias score. This amplifies the noise caused by mistakes of the NER system and makes their results inconclusive.

Assuming unbiased NER systems and balanced data, the baselines should achieve a bias score very close to 1. The results of all baselines on the GAP test set is reported in the left part of Table 1, where we see that most of the bias scores strongly differ from 1. In the next section, we show that imbalanced data are the reason behind this. Notice that the acc-Bias score of a model is usually further from 1 than its $F_1$-Bias score. These results empirically support our intuition that $F_1$-Bias is less representative than acc-Bias, as noise from negative candidates makes $F_1$-Bias less sensitive. Thus, the accuracy-based bias metric is more appropriate than its $F_1$-score counterpart.

![Table 1: Performance and bias metrics on baseline systems on GAP, implemented with two different NER systems as well as the ground-truth personal names. The reported performance of the random classifier is obtained by averaging the performance over 10,000 repetitions. If the evaluated baseline is expected to achieve a score of 1 on some metric due to balancing, the score is written in italics. Note that deviations can happen when NER-system extractions are incorrect.](https://cloud.google.com/natural-language/)

| baseline            | $F_1$ Accuracy | $F_1$-Bias acc-Bias | W-Bias | W$_{num}$-Bias | W$_{dist}$-Bias | W$_t$-Bias |
|---------------------|----------------|---------------------|--------|----------------|----------------|------------|
| Ground-Truth Random | 0.305          | 0.224               | 0.884  | 0.849          | 1.000          | 0.995      | 0.899      | 1.000      |
| Spacy-lg Random     | 0.286          | 0.211               | 0.904  | 0.870          | 0.975          | 0.980      | 0.905      | 0.984      |
| Google-NER Random   | 0.295          | 0.218               | 0.937  | 0.907          | 1.019          | 1.021      | 0.949      | 1.020      |
| Ground-Truth Dist-1 | 0.463          | 0.412               | 0.850  | 0.776          | 1.000          | 0.804      | 1.000      | 1.000      |
| Spacy-lg Dist-1     | 0.423          | 0.375               | 0.887  | 0.816          | 1.015          | 0.824      | 1.029      | 1.018      |
| Google-NER Dist-1   | 0.446          | 0.399               | 0.875  | 0.799          | 0.986          | 0.793      | 1.016      | 0.994      |
| Ground-Truth Dist-2 | 0.353          | 0.310               | 0.923  | 0.882          | 1.000          | 0.920      | 1.000      | 1.000      |
| Spacy-lg Dist-2     | 0.319          | 0.263               | 0.917  | 0.907          | 0.962          | 0.932      | 0.977      | 0.968      |
| Google-NER Dist-2   | 0.354          | 0.309               | 0.946  | 0.915          | 1.000          | 0.983      | 1.001      | 1.026      |
| Ground-Truth Dist-3 | 0.228          | 0.156               | 1.270  | 1.347          | 1.000          | 1.206      | 1.010      | 1.007      |
| Spacy-lg Dist-3     | 0.205          | 0.134               | 1.490  | 1.585          | 1.118          | 1.494      | 1.152      | 1.200      |
| Google-NER Dist-3   | 0.219          | 0.150               | 1.312  | 1.426          | 1.154          | 1.368      | 1.111      | 1.116      |
3.3 Analysis of GAP

Our analysis of the manually annotated spans of personal names shows that masculine examples contain 5.55 personal names on average (standard deviation 3.18), while feminine examples contain 6.30 names on average (standard deviation 3.44). This confirms the hypothesis about imbalances in the data and explains why the Ground-Truth random baseline achieved a bias score different from 1. The full distribution of the number of names per sentence is given in Appendix E.

Secondly, we sort all annotated personal names in each sentence by distance to the pronoun in the same way as done by the Dist-k baselines. We find the position of the correct candidate on this ordered list. The average position of the correct candidate in the masculine subset is 1.86 (standard deviation 1.19) candidates away from the pronoun, while the average position in the feminine subset is 2.32 (standard deviation 1.54) candidates away from the pronoun, potentially explaining the bias scores of the Dist-k baselines. Examples with no correct candidate are not considered in this statistic. The full distribution is given in Appendix E.

3.4 Weighting GAP

Using our manual annotations of personal names, let $N_k$ be the set of all examples with exactly $k$ personal names, and let $D_k$ be the set of all examples where the correct candidate is the $k$-th closest candidate to the pronoun. The $N_k$ and $D_k$ sets form the sets that we generically denoted as $S_1, \ldots, S_m$ in Section 2. Thus, the sets $N_k$ and $D_k$ are used as input to our balancing method to obtain a linear program, which we solve with the LINPROG optimization tool from Matlab, version R2019b.

We name the obtained weighted bias metric $W$-Bias. We highlight that the introduced constraints are not a guarantee that $W$-Bias is completely balanced, as other imbalances in the data may exist. However, given Theorem 2.1, known imbalances have been balanced out, while introducing the least noise possible, making the introduced metric preferred over the existing one, i.e., no weighting. A visualization of the weights is in Section C.

To assess the introduced weights, we evaluate the baselines on the newly introduced W-Bias metric. To confirm that our method does not introduce noise relative to unidentified properties, we perform two ablation experiments. In the first one, we ignore the distance property, while in the second experiment, we ignore the number of candidates. To this end, we introduce two more bias metrics: $W_{\text{num}}$-Bias and $W_{\text{dist}}$-Bias. In $W_{\text{num}}$-Bias, the sets $D_k, k \in \mathbb{N}$, were not included as the input to the balancing procedure. $W_{\text{num}}$-Bias is only balanced with respect to the number of names per sentence. On the other hand, $W_{\text{dist}}$-Bias does not include the sets $N_k, k \in \mathbb{N}$, meaning that it is only balanced relative to the distance between the pronoun and the correct answer. We show that, for random baselines, the following holds: $|1 - W_{\text{Bias}}| \leq |1 - W_{\text{dist}}| \leq |1 - \text{acc-Bias}|$, that is, balancing relative to distance does not exacerbate bias scores of random baselines, and that additional balancing relative to number of names further decreases its distance to unbiased score (of 1). Similarly, we show that for Dist-k baselines, $|1 - W_{\text{Bias}}| \leq |1 - W_{\text{num}}| \leq |1 - \text{acc-Bias}|$.

The results are reported in Table 1. In the columns that correspond to $W_{\text{num}}$-Bias and $W_{\text{dist}}$-Bias, numbers in italics are expected to be similar to the numbers predicted by W-Bias. We see that the inequations in the previous paragraph hold for all baselines, showing that our weights indeed do not exacerbate the bias of unidentified properties. Moreover, we see that the W-Bias scores achieved by the baselines are consistently closer to 1 than their acc-Bias scores, confirming that the introduced weights balance the bias metric. In particular, the W-Bias score of Ground-Truth baselines is equal to 1, i.e., unbiased. We note that the minimal deviation from 1 of the Ground-Truth W-bias score for the Dist-3 baseline is a consequence of a disagreement between our span annotations with the spans of gold labels. Bias scores of the Dist-2 and Dist-3 baselines implemented with NER systems are subject to larger deviations that happen because these baselines are more sensitive to disagreement between the NER system and our annotations of the name spans.

We note that weighting with respect to one of the imbalances sometimes helped balancing the baseline that was affected by the other. For example, balancing the number of names per sentence ($W_{\text{num}}$-Bias) resulted in improved bias scores of all Ground-Truth Distance baselines. This implies that there exists a correlation between the number of personal names in the sentence and the distance between the pronoun and the correct candidate in the GAP test set.
An analysis of W-Bias weights shows that the distribution of weights contains some outliers, that is, examples with unusually large weights, see Appendix C for a more detailed discussion. Ten examples with largest weights have a weight average of 6.29 (the average weight overall is 1.0). These examples all come from examples in $D_k$ and $N_k$ with large $k$, because these sets are often highly gender-imbalanced, as discussed in Appendix B. While this is theoretically correct, it may be undesirable, as it means that few out-of-distribution examples carry a lot of weight in the final score, which could introduce noise.

We show that such large weights can be avoided by removing highly-imbalanced subsets of the data. We introduce a trimmed $W$-score, called $W_t$-score. Examples with more than 15 personal names and examples where the correct candidate is the $k$-th closest for $k \geq 5$ are removed from this score, reducing the size of the dataset to 1670 examples (83.5% of the original size). Numbers $k \geq 5$ and 15 personal names were selected manually by consulting figures which can be found in Appendix B. The rest of the examples are assigned new weights with the introduced method. Top ten weights of $W_t$-score have a weight average of 3.3, strongly reducing the problem of outliers. Comparing $W_t$-bias with $W$-bias in Table 1 shows that such outliers mainly affected Dist-2 and Dist-3 baselines.

### 3.5 Evaluation of Bias in Existing Coreference Models

|      | $F_1$ | $F_1$-Bias | acc-Bias | W-Bias | $W_t$-Bias |
|------|-------|------------|----------|--------|-----------|
| 1Bert         | 0.500 | 0.88       | 0.86     | 0.85   | 0.87      |
| 1Bert_WikiCREM| 0.590 | 0.95       | 0.93     | 0.90   | 0.92      |
| 1Bert_GAP     | 0.752 | 0.99       | 0.97     | 0.96   | 0.97      |
| 1Bert_DPR     | 0.612 | 1.00       | 0.96     | 0.94   | 0.96      |
| 1Bert_ALL     | 0.760 | 1.03       | 1.03     | 1.03   | 1.04      |
| 1Bert_GAP_DPR | 0.704 | 1.01       | 1.00     | 1.00   | 1.00      |
| 1Bert_WikiCREM_GAP | 0.778 | 1.01 | 1.00     | 1.00   | 1.01      |
| 1Bert_WikiCREM_DPR | 0.646 | 0.99 | 0.98     | 0.97   | 0.96      |
| 1Bert_WikiCREM_ALL | 0.783 | 1.02 | 1.01     | 1.00   | 1.01      |
| 2Bert_BASE    | 0.824 | 0.97       | 0.97     | 0.96   | 0.96      |
| 2Bert_LARGE   | 0.856 | 0.97       | 0.96     | 0.96   | 0.97      |
| 3SpanBERT_BASE| 0.855 | 0.96       | 0.95     | 0.95   | 0.95      |
| 3SpanBERT_LARGE| 0.877 | 0.95 | 0.94     | 0.93   | 0.93      |
| 4E2E          | 0.733 | 0.93       | 0.92     | 0.92   | 0.91      |
| 4E2E_ADV      | 0.747 | 0.93       | 0.91     | 0.93   | 0.90      |
| 8RefReader    | 0.794 | 0.96       | 0.95     | 0.97   | 0.97      |

Table 2: [Kocijan et al. 2019]; [Joshi et al. 2019a]; [Joshi et al. 2019b]; [Lee et al. 2018]; [Subramanian and Roth 2019]; [Liu et al. 2019]. We used publicly shared code and models for all models, except for Referential Reader [Liu et al. 2019], where code was not publicly available at the time. Instead, the evaluation was performed on the results provided by the authors. The numbers differ from the paper, because the authors averaged results over several seeds, but only shared one version. The results from [Joshi et al. 2019b] differ from the ones in the paper, as the author shared a different checkpoint. Evaluation of several state-of-the-art models for coreference resolution on GAP, with several bias scores reported.

Having shown that the introduced measure strongly reduces the impact of the observed imbalances in the dataset, we re-evaluate recently released models for coreference resolution. Following [Webster et al. 2018], we only consider systems that detect name spans for inference automatically and access labelled spans only to output predictions. We thus do not re-evaluate models that were submitted as part of the Kaggle competition on the GAP dataset, because they do not conform to this norm [Webster et al. 2019]. The results are reported in Table 2. A description of how statistical significance of changes in $W_t$-Bias can be computed can be found Appendix D.

Comparing acc-Bias and W-Bias, we can see that only a few models change their bias score visibly, indicating that not all models were equally affected by the observed imbalances. While we cannot directly compare these bias scores with the original $F_1$-based bias metric, we hypothesize that the imbalances in the data also affected that score. Comparing W-Bias and $W_t$-Bias shows that most of the models were minimally affected by the outliers in the weights.
We observe that the better performing models tend not to change their bias scores significantly. We hypothesize that they are less affected by the observed imbalances in the data distribution. At the same time, a larger denominator (female score) in the bias formula results in a smaller absolute difference. Similarly, we can see that RNN-based models (models 4,5,6) change their scores more than transformer-based models (models 1,2,3), implying that RNN-based models were more affected by the number of candidates and the distance between the correct candidate and the pronoun than transformer-based models.

### 4 Related Work

An increasing amount of work has recently been done on fairness both in NLP and machine learning in general. NLP datasets for bias detection mainly detect gender bias in coreference resolution [Zhao et al. 2018], [Rudinger et al. 2018], [Webster et al. 2018], however, other types of bias-detection dataset exist. For example, Etc. Kiritchenko and Mohammad (2018) is focused on sentiment analysis and additionally measures racial bias. All listed datasets other than GAP are constructed artificially, often from hand-written templates. Thus, they do not suffer from the irregularities observed in GAP, however, they also do not reflect the bias on the real-world data. We refer to [Mehrabi et al. 2019] for an overview of fairness in machine learning, and to [Sun et al. 2019] for a more specific review on biases in NLP.

Quite some work has been done in debiasing or otherwise balancing training data. Zhao et al. [2018] show how debiased word embeddings and swapping gender of pronouns and antecedents can be used to reduce the gender bias of coreference models. Kamiran and Calders [2012] and Chandrasekaran and Kan [2018] weight training examples to remove bias from the training data, the latter also using linear programming, similarly to us. However, they only balance the data with respect to a single property, while our method works for several. Sakaguchi et al. [2019] aim to reduce the bias in coreference resolution caused by annotation artifacts both in the training and test data. Their main aim is to remove the systemic bias in the dataset that could give away unintended cues on the correct answers. We highlight that their work concerns a different type of bias, as their goal is to prevent any model from achieving a high performance due to spurious correlations in the dataset, rather than to reduce any type of discrimination between different candidates.

### 5 Summary and Outlook

In this work, we introduced a test-set weighting method to remove undesired imbalances in bias-measuring datasets, without exacerbating other potentially undesired patterns.

We demonstrated the method on the GAP test set, which contained such undesired irregularities. We annotated the dataset with spans of all personal names and introduced the bias metrics W-Bias and W_t-Bias that balance out the observed irregularities. While there is no guarantee that these scores balance out all data irregularities in GAP, we showed that they balance out the ones that we are aware of. We encourage research to use our introduced metrics to measure the bias of coreference models on GAP. Among the two introduced scores, we recommend W_t because its weights contain fewer outliers, and the chance of undesired deviations in the score is smaller.

The introduced weighting method can easily be applied to other datasets. A room for improvement of the method is to remove the need to identify the biases in the data, however, this step is common to existing methods that deal with bias. It is not unreasonable to expect that the existence of bias has to be noticed and demonstrated before one can start planning the debiasing. This already satisfies the prerequisites to use the introduced method. Manual annotation of examples like the one in our work is not always necessary, as automatic tools (e.g., NER systems) can be used. However, manual annotation likely ensures the high quality of the test data.

This work addresses an important problem of real-world bias metrics and opens up several questions. Is it possible to balance out unobserved irregularities in the data that could negatively affect a model? Can the weighting method be extended to non-linear metrics, such as $F_1$-score? Can we construct better bias datasets with real-world examples that contain fewer patterns that could affect a bias measure in undesired ways? We encourage research into the direction of detecting bias on real-world data. Finally, this work encourages deeper reflections into designing bias-testing protocols.
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References

Lauren Ackerman. Syntactic and cognitive issues in investigating gendered coreference. *Glossa: a journal of general linguistics*, 4(1):117, 2019.

Muthu Kumar Chandrasekaran and Min-Yen Kan. Countering position bias in instructor interventions in MOOC discussion forums. In *Proceedings of the 5th Workshop on Natural Language Processing Techniques for Educational Applications*, pages 135–142, July 2018.

Ali Emami, Paul Trichelair, Adam Trischler, Kaheer Suleman, Hannes Schulz, and Jackie Chi Kit Cheung. The KnowRef coreference corpus: Removing gender and number cues for difficult pronominal anaphora resolution. In *Proceedings of the 57th Annual Meeting of the Association for Computational Linguistics*, pages 3952–3961, July 2019.

Mandar Joshi, Danqi Chen, Yinhan Liu, Daniel S. Weld, Luke Zettlemoyer, and Omer Levy. SpanBERT: Improving pre-training by representing and predicting spans. *arXiv preprint arXiv:1907.10529*, 2019a.

Mandar Joshi, Omer Levy, Luke Zettlemoyer, and Daniel Weld. BERT for coreference resolution: Baselines and analysis. In *Proceedings of the 2019 Conference on Empirical Methods in Natural Language Processing and the 9th International Joint Conference on Natural Language Processing (EMNLP-IJCNLP)*, pages 5803–5808, November 2019b.

Faisal Kamiran and Toon Calders. Data preprocessing techniques for classification without discrimination. *Knowledge and Information Systems*, 33(1):1–33, 2012.

Svetlana Kiritchenko and Saif Mohammad. Examining gender and race bias in two hundred sentiment analysis systems. In *Proceedings of the 7th Joint Conference on Lexical and Computational Semantics*, pages 43–53, June 2018.

Vid Kocijan, Oana-Maria Camburu, Ana-Maria Cretu, Yordan Yordanov, Phil Blunsom, and Thomas Lukasiewicz. WikiCREM: A large unsupervised corpus for coreference resolution. In *Proceedings of the 2019 Conference on Empirical Methods in Natural Language Processing and the 9th International Joint Conference on Natural Language Processing (EMNLP-IJCNLP)*, pages 4303–4312, November 2019.

Kenton Lee, Luheng He, and Luke Zettlemoyer. Higher-order coreference resolution with coarse-to-fine inference. In *Proceedings of the 2018 Conference of the North American Chapter of the Association for Computational Linguistics: Human Language Technologies, Volume 2 (Short Papers)*, pages 687–692, June 2018.

Fei Liu, Luke Zettlemoyer, and Jacob Eisenstein. The referential reader: A recurrent entity network for anaphora resolution. *CoRR*, abs/1902.01541, 2019.

Ninareh Mehrabi, Fred Morstatter, Nripsuta Saxena, Kristina Lerman, and Aram Galstyan. A survey on bias and fairness in machine learning. *arXiv preprint arXiv:1908.09635*, 2019.

Altarf Rahman and Vincent Ng. Resolving complex cases of definite pronouns: The Winograd Schema Challenge. In *Proceedings of the 2012 Joint Conference on Empirical Methods in Natural Language Processing and Computational Natural Language Learning*, pages 777–789, July 2012.

Rachel Rudinger, Jason Naradowsky, Brian Leonard, and Benjamin Van Durme. Gender bias in coreference resolution. In *Proceedings of the 2018 Conference of the North American Chapter of the Association for Computational Linguistics: Human Language Technologies*, June 2018.
Keisuke Sakaguchi, Ronan Le Bras, Chandra Bhagavatula, and Yejin Choi. WINOGRANDE: An adversarial Winograd Schema Challenge at scale. *CoRR*, abs/1907.10641, 2019.

Sanjay Subramanian and Dan Roth. Improving generalization in coreference resolution via adversarial training. In *Proceedings of the Joint Conference on Lexical and Computational Semantics*, 6 2019.

Tony Sun, Andrew Gaut, Shirlyn Tang, Yuxin Huang, Mai ElSherief, Jieyu Zhao, Diba Mirza, Elizabeth Belding, Kai-Wei Chang, and William Yang Wang. Mitigating gender bias in natural language processing: Literature review. In *Proceedings of the 57th Annual Meeting of the Association for Computational Linguistics*, pages 1630–1640, July 2019.

Kellie Webster, Marta Recasens, Vera Axelrod, and Jason Baldridge. Mind the GAP: A balanced corpus of gendered ambiguous pronouns. *Transactions of the Association for Computational Linguistics*, 6:605–617, 2018.

Kellie Webster, Marta R. Costa-jussà, Christian Hardmeier, and Will Radford. Gendered Ambiguous Pronoun (GAP) Shared Task at the Gender Bias in NLP Workshop 2019. In *Proceedings of the 1st Workshop on Gender Bias in Natural Language Processing*, pages 1–7, August 2019.

Alexander Yeh. More accurate tests for the statistical significance of result differences. In *COLING 2000 Volume 2: The 18th International Conference on Computational Linguistics*, 2000.

Jieyu Zhao, Tianlu Wang, Mark Yatskar, Vicente Ordonez, and Kai-Wei Chang. Gender bias in coreference resolution: Evaluation and debiasing methods. In *Proceedings of the 2018 Conference of the North American Chapter of the Association for Computational Linguistics: Human Language Technologies, Volume 2 (Short Papers)*, pages 15–20, June 2018.
A Proof of Theorem 2.1

We derive the criterion function for the set $A$, as the derivation for set $B$ is analogous. We denote

$$T_C := |\text{Acc}_W^A(C \cap A) - \text{Acc}^A(C \cap A)|.$$ 

We first simplify $T_C$, piece by piece:

$$\text{Acc}^A(C \cap A) = \frac{|C \cap A|}{|A|},$$

$$\text{Acc}_W^A(C \cap A) = \frac{2}{n} \sum_{x_i \in C \cap A} w_i.$$

Combining this with the formula for $T_C$, we get:

$$T_C = \frac{2}{n} \left( \sum_{x_i \in C \cap A} w_i \right) - \frac{|C \cap A|}{|A|}$$

$$= \frac{2}{n} \sum_{x_i \in C \cap A} \left( w_i - \frac{n}{2|A|} \right).$$

To minimize $|T_C|$, we have to minimize

$$\left| \sum_{x_i \in C \cap A} \left( w_i - \frac{n}{2|A|} \right) \right|.$$ 

To minimize the upper bound, we take a look at the scenarios that give the largest value of $|T_C|$. Let $w_{A(i)}^*$ be the $i$-th smallest weight corresponding to an example in $A$. We use the constant $\lambda := \frac{n}{2|A|}$.

The following properties hold:

$$\sum_{i=1}^{\lfloor A \rfloor} w_{A(i)} = \sum_{x_i \in A} w_i = \frac{n}{2}$$

$$\sum_{i=1}^{\lfloor A \rfloor} (w_{A(i)} - \lambda) = \sum_{i=1}^{\lfloor A \rfloor} (w_{A(i)} - \frac{n}{2|A|})$$

$$= \sum_{x_i \in A} w_i - \frac{n}{2} = 0.$$ 

In the scenario where $T_C$ is maximal, $C \cap A$ will include $|C \cap A|$ examples with either largest or smallest weights. Let $k := |C \cap A|$, and let us first take a look at the case where examples with largest $k$ weights are in $C \cap A$. To minimize all such cases, we aim to minimize the following term:

$$(3) \sum_{k=1}^{\lceil A \rceil} \sum_{i=\frac{A}{2}+k+1}^{\lfloor A \rfloor} (w_{A(i)} - \lambda) = \sum_{i=1}^{\lceil A \rceil} (w_{A(i)} - \lambda)i.$$
On the other hand, we aim to maximize the opposite case, i.e., when examples with the smallest $k$ weights are in $C \cap A$. The objective that we aim to maximize can be written as follows:

$$\sum_{k=1}^{\left|A\right|} \sum_{i=1}^{k} (w_{A(i)} - \lambda) =$$

$$= - \sum_{k=1}^{\left|A\right|} \sum_{i=k+1}^{\left|A\right|} (w_{A(i)} - \lambda)$$

$$= - \sum_{i=1}^{\left|A\right|} (w_{A(i)} - \lambda)i + \sum_{i=1}^{\left|A\right|} (w_{A(i)} - \lambda)$$

$$= - \sum_{i=1}^{\left|A\right|} (w_{A(i)} - \lambda)i + 0$$

$$= - \sum_{i=1}^{\left|A\right|} (w_{A(i)} - \lambda)i .$$

We can see that maximizing this term is equivalent to minimizing term (3). Minimizing term (3) is therefore sufficient. We further simplify it:

$$\sum_{i=1}^{\left|A\right|} (w_{A(i)} - \lambda)i =$$

$$= \sum_{x_i, x_j \in A \atop i > j} \max(w_i - \lambda, w_j - \lambda) + \sum_{x_i \in A} (w_i - \lambda)$$

$$= \sum_{x_i, x_j \in A \atop i > j} \max(w_i - \lambda, w_j - \lambda) + 0$$

$$= \sum_{x_i, x_j \in A \atop i > j} \max(w_i, w_j) - \lambda \cdot \frac{|A||A| - 1}{2} .$$

Since the second part of the term is constant, we aim to minimize the following objective:

$$\sum_{x_i, x_j \in A \atop i > j} \max(w_i, w_j) .$$

In the same way, the objective for the examples in set $B$ can be computed. Summing them up, we obtain the following objective:

$$\min \sum_{x_i, x_j \in A \atop i > j} \max(w_i, w_j) + \sum_{x_i, x_j \in B \atop i > j} \max(w_i, w_j) .$$

### B Analysis of the GAP Data Distribution

This section contains information on the distribution of the GAP test set. Fig.[1] shows the distribution of the number of names per sentence in the masculine and the feminine subset of the test set. The difference between the masculine and the feminine subset are visible, with masculine examples usually containing fewer names per sentence in comparison to feminine examples. This histogram clearly shows why the random baseline did not achieve a bias score around $1$.

In Fig.[2] we can see how often the correct entity is the closest one to the pronoun, how often it is the second closest, third closest, etc. We can observe a similar pattern, with the closest and second closest candidate being correct more often in the masculine subset. The distribution is the source of a seemingly biased performance of the Dist-$k$ baselines.
We can see that sets at the tails of both two graphs can be highly gender-imbalanced. As discussed in Sections 3.4 and C, the weighting method counteracts these imbalances by assigning large weights to the weights in the under-represented class.

We additionally visualize the dataset after the trimming that was done as part of the $W_t$-Bias score. The distribution of the number of names and how close the correct candidate is after trimming are reported in Figs. 3 and 4. The trimmed dataset contains fewer highly-imbalanced subsets. We did not conduct any further trimming of the dataset to avoid a further reduction of its size.

Figure 1: Histogram showing the number of personal names per sentence in the GAP dataset. The X-axis shows the number of names in the sentence, and the Y-axis the number of sentences with the corresponding number of personal names. The blue and green columns show the data for masculine and feminine examples, respectively.

Figure 2: Histogram showing how often the correct entity is the closest, second closest, third closest, etc. entity to the pronoun in the GAP dataset. The blue and green columns show the data for masculine and feminine examples, respectively.

Figure 3: Histogram showing the number of personal names per sentence in the GAP dataset after the trimming conducted for the $W_t$-Bias score. The blue and green columns show the data for masculine and feminine examples, respectively.

Figure 4: Histogram showing how often the correct entity is the closest, second closest, third closest, etc. entity to the pronoun in the GAP dataset after the trimming conducted for the $W_t$-Bias score. The blue and green columns show the data for masculine and feminine examples, respectively.

C Visualization of Weights

A visualisation of W-Bias weights of positive examples is shown in Figure 5. The visualisation confirms that weights gravitate around 1 despite the constraints. Moreover, fewer than 1% of the weights are set to 0. A manual investigation into these examples shows that they are often very long text spans with long lists of names, such as family trees or cast lists. Several of them are not grammatically correct sentences, but rather lists from Wikipedia that were not removed during the annotation.

There are 9 weights larger than 4.0 that are not pictured: 4.84, 4.97, 4.97, 5.43, 6.41, 7.05, 7.05, 9.20, and 9.72, all of them corresponding to male examples. Their weights are large, because they
Figure 5: Histogram showing the distribution of W-Bias weights, split into intervals of size 0.1. The blue parts of the columns correspond to masculine examples, while the green parts correspond to feminine examples. The weights are centered around 1. Nine largest weights are not included in the histogram, as they have values over 4.

A visualization of W1-Bias weights on Figure 6 shows that trimming largely solves this problem, as the largest few weights now carry much less weight than before, as discussed in Section 3.4. Moreover, there are no examples with weight 0. Female weights are slightly larger on average, because there are more male (865) than female (805) examples in the trimmed dataset. We do not conduct any additional trimming to avoid decreasing the size of the dataset further.

D Statistical significance of bias metrics

We used the randomization test [2000] to compare the W1-Bias scores of a few models, re-evaluated in Section 3.5. E.g., the difference between BERT and BERT_GAP is significant ($p = 0.024$), as is the difference between BERT_WikiCREM and BERT_WikiCREM_GAP ($p = 0.017$). Finetuning BERT on GAP thus seems to significantly increase its bias score. On the other hand, the difference of the E2E and E2E_ADV models is not significant ($p = 0.364$), implying that the seemingly negative impact of adversarial sampling could be a coincidence.