Abstract—We present the design of a motion planning algorithm that ensures safety for an autonomous vehicle. In particular, we consider a multimodal distribution over uncertainties; for example, the uncertain predictions of future trajectories of surrounding vehicles reflect discrete decisions, such as turning or going straight at intersections. We develop a computationally efficient, scenario-based approach that solves the motion planning problem with high confidence given a quantifiable number of samples from the multimodal distribution. Our approach is based on two preprocessing steps, which 1) separate the samples into distinct clusters and 2) compute a bounding polytope for each cluster. Then, we rewrite the motion planning problem approximately as a mixed-integer problem using the polytopes. We demonstrate via simulation on the nuScenes dataset that our approach ensures safety with high probability in the presence of multimodal uncertainties, and is computationally more efficient and less conservative than a conventional scenario approach.

Index Terms—Stochastic optimal control; Autonomous vehicles.

I. INTRODUCTION

SAFE motion planning in uncertain environments has been an active research area with consideration of various sources of uncertainties. Disturbances to dynamic models, e.g., wind disturbances, are the commonly studied source of uncertainties [1], [2]. Uncertain specifications, e.g., uncertain obstacle locations, are another practical source [3]–[5]. Several past approaches assume a Gaussian distribution over uncertainties [1], [2], [5] or represent the distribution via a finite number of samples without any assumption on it [3], [4].

In this paper, we focus on a safe motion planning algorithm that can efficiently handle multimodal uncertainties. We are motivated by an autonomous driving example, where future trajectories of other vehicles (OVs) are forecasted by a deep neural network (DNN), and a motion planner computes a trajectory of the ego vehicle (EV) that avoids any collision with the forecasts (See Fig. 1). A noteworthy feature of most forecasting methods (e.g., [6]) is that the distribution over future trajectories is multimodal. In other words, predicted trajectories reflect different high-level decisions of OVs, such as braking/accelerating or going-straight/turning. Given prediction samples from the multimodal distribution, our goal is to design a motion planning algorithm that guarantees safety with high probability. A similar problem is studied in [7], but focuses on the design of a forecasting method to facilitate the integration of forecasting into control.

Stochastic control with multimodal distributions is a more challenging problem than that with unimodal distributions. For example, if an uncertain parameter has a Gaussian distribution, a chance constraint can be reformulated as a convex constraint in terms of the quantile function [8]. However, with a mixture of two Gaussian distributions, such a convex reformulation is not possible due to the lack of a closed form for its quantile function. To our knowledge, there is no tractable way to solve general chance-constrained problems with multimodal distributions. In [9], a branch-and-bound approach to solve chance-constrained problems with a mixture of Gaussian distributions is presented, but the approach is not tractable for large problems and thus not applicable to real-time implementation.

We formulate the motion planning problem as a chance-constrained problem, and present an approximate yet efficient method to solve the problem when the uncertainty has a multimodal distribution. Specifically, we use the scenario approach [10] to characterize the number of samples \( N \) required to solve the chance-constrained problem with high confidence. We show that the conventional scenario approach [10] is only applicable to an approximation of the chance-constrained problem. Furthermore, it yields a conservative solution when the uncertainty has a multimodal distribution, and is computationally demanding as the number of mixed-integer constraints depends on the number of samples \( N \). Our approach to addressing these issues uses two preprocessing steps: First, we separate samples into distinct clusters to explicitly handle the
multimodality, and next, we compute a bounding polytope for each cluster. With the preprocessing steps, we can rewrite the chance-constrained problem approximately as a deterministic problem, whose computational complexity does not depend on $N$. Our approach is inspired by the preprocessing-based solution to chance-constrained problems [11]. We extend their technique to address the specific challenges of motion planning for autonomous driving: namely, non-convex collision-avoidance constraints and multimodal uncertainties.

In summary, the contributions of this paper are:

- We show that applying the conventional scenario approach [10] to a motion planning problem yields conservative results when the uncertainty has a multimodal distribution and is computationally expensive.
- We present a computationally efficient method to solve the motion planning problem with high confidence.
- We test our approach with a state-of-the-art forecasting DNN on the nuScenes dataset [12] and validate that it ensures a desired level of safety, is computationally efficient, and is less conservative than the conventional scenario approach.

**Notation:** We denote by $\mathbb{R}^n$ the set of $n$-dimensional real vectors, and $\mathbb{Z}, \mathbb{Z}_{a:b}$ by the sets of integers and integers bounded by $a \in \mathbb{Z}$ and $b \in \mathbb{Z}$, i.e., \{a, a + 1, \ldots, b\}, respectively. Let $x_t \in \mathcal{X} \subseteq \mathbb{R}^{n_x}$, $u_t \in \mathcal{U} \subseteq \mathbb{R}^{n_u}$, $y_t \in \mathcal{Y} \subseteq \mathbb{R}^{n_y}$ be the dynamical state, input, and output of the ego vehicle at time step $t$, respectively. A sequence of inputs over $t \in \mathbb{Z}_{0:T-1}$ is denoted by the bold face $\mathbf{u} := (u_0, \ldots, u_{T-1})$ in the space $\mathcal{U} = \mathcal{U} \times \cdots \times \mathcal{U}$. Capital roman-type letters, such as $X$, indicate random variables.

## II. Problem Setup

We aim to compute the input sequence $\mathbf{u} \in \mathcal{U}$ over a planning horizon of length $T$ that solves the following motion planning problem.

\[
\begin{align*}
\min_{\mathbf{u} \in \mathcal{U}} & \quad J(\mathbf{u}) \quad (1a) \\
\text{s.t.} & \quad x_{t+1} = f_t(x_t, u_t), \quad t \in \mathbb{Z}_{0:T-1}, \quad (1b) \\
& \quad y_t = h_t(x_t), \quad t \in \mathcal{X}_{1:T}, \quad (1c) \\
& \quad \Pr \left( \bigwedge_{t=1}^{T} y_t \notin \mathcal{O}(x_{t:o}) \right) \geq 1 - \epsilon. \quad (1d)
\end{align*}
\]

Here, $J : \mathcal{U} \to \mathbb{R}$ is a convex objective function that guides the EV toward goal states, (1b) is the discrete-time dynamical model of the EV, (1c) is the output, and (1d) is a chance constraint that encodes the avoidance of collisions. In constraint (1d), $x_{t:o} \in \mathbb{R}^{n_x}$ is the predicted state of OV $o$ at time $t$, and $\mathcal{O}(x_{t:o}) \subseteq \mathbb{R}^{n_y}$ is an obstacle set that the EV's output $y_t$ should avoid. Since $x_{t:o}$ is a random variable, we impose the probability of non-collision to be no smaller than $1 - \epsilon$, where $\epsilon \in (0, 1)$ indicates a risk level. The number of OVs around the EV is $O$.

Consider that the dynamical model is linear time-varying, that is, at each time step $t$, $f_t$ and $h_t$ in (1b) and (1c) are linear functions of $x_t$ and $u_t$. Also, let $\mathcal{O}(x_{t:o})$ be a polytope expressed by

\[
\mathcal{O}(x_{t:o}) := \{ y \in \mathbb{R}^{n_y} : A(x_{t:o})y < b(x_{t:o}) \},
\]

which is the intersection of $L$ halfspaces defined by nonlinear continuous functions $A : \mathbb{R}^{n_x} \to \mathbb{R}^{L \times n_y}$ and $b : \mathbb{R}^{n_x} \to \mathbb{R}^L$. To encode constraints of polytopic obstacle avoidance, it is common to use the big-$M$ method [13] where binary variables are introduced to indicate whether the inequalities $A(x_{t:o})y < b(x_{t:o})$ in (2) are satisfied. Specifically, we can introduce a binary variable $z_{j,t:o}(x_{t:o}) \in \{0, 1\}$ as a function of $x_{t:o}$ and rewrite $y_t \notin \mathcal{O}(x_{t:o})$ as

\[
A_j(x_{t:o})y_t + M \left(1 - z_{j,t:o}(x_{t:o})\right) \geq b_j(x_{t:o}), \quad \forall j \in \mathbb{Z}_{1:L}, \quad (3)
\]

and $\sum_{j=1}^{L} z_{j,t:o}(x_{t:o}) \geq 1$, where $M \in \mathbb{R}$ is a sufficiently large number, and $A_j$ and $b_j$ are the $j$-th row of $A$ and $b$, respectively. Here, $z_{j,t:o}(x_{t:o}) = 1$ indicates that one of the inequalities in (2) is violated (i.e., $A_j(x_{t:o})y_t \geq b_j(x_{t:o})$), which implies that output $y_t$ is outside $\mathcal{O}(x_{t:o})$.

Next, we explain forecasting algorithms that generate $\mathbf{X} = (x_{t:o} : t \in \mathbb{Z}_{1:T}, o \in \mathbb{Z}_{1:o})$, and discuss our approach to solving problem (1) based on the forecasts.

### A. Forecasting

Forecasting algorithms generate future trajectories of OVs based on context $\phi$, which is often a motion history of each OV and current road structure perceived by sensors, e.g., cameras and lidars. Given a training dataset of ground-truth future trajectories and contexts, i.e., $(\mathbf{X}^t, \phi^t)$, the goal of recent forecasting DNNs (e.g., [6]) is to learn the conditional probability distribution $p(\mathbf{X} | \phi)$, so that at test time, they can extract arbitrarily many predictions from it.

Since we focus on the design of a motion planner, we assume that a forecasting DNN is well-trained with a sufficient amount of training data and accurately predicts possible future trajectories of OVs. This is to avoid out-of-distribution scenarios (i.e., scenarios that have not been observed during training), which may cause incorrect forecasts. The relaxation of this assumption leads to a different research problem actively studied in machine learning communities [14], and remains as future work.

### B. Scenario Approach

A common approach to chance-constrained problems is the scenario approach [15], which represents the uncertainty distribution via a finite number of samples. When the distribution is modeled by a DNN, we can query the DNN to generate samples from the distribution. While [15] is applicable only to convex problems, the work [10] extends it to mixed-integer problems. However, the reformulation (3) does not fit in the formulation from [10] because binary decision variables in (3) depend on the uncertain parameter $x_{t:o}$. Thus we instead consider the following approximation of problem (1).

\[
\begin{align*}
\min_{\mathbf{u} \in \mathcal{U}, z \in \mathbb{Z}} & \quad J(\mathbf{u}) \quad (4a) \\
\text{s.t.} & \quad \bigwedge_{j=1}^{L} A_j(x_{t:o})y_t + M \left(1 - z_{j,t:o}(x_{t:o})\right) \geq b_j(x_{t:o}), \quad \forall j \in \mathbb{Z}_{1:L}, \quad (4b) \\
& \quad \Pr \left( \bigwedge_{j=1}^{L} A_j(x_{t:o})y_t + M \left(1 - z_{j,t:o}(x_{t:o})\right) \geq b_j(x_{t:o}) \right) \geq 1 - \epsilon, \quad (4c)
\end{align*}
\]
where \( z \) is a sequence of binary variables \( z_{j,t,o} \in \{0, 1\} \) for \( j \in Z_{1:L}, t \in Z_{1:T}, o \in Z_{1:O}, \) which are independent of \( X_{t,o} \) in contrast to \( c \), and \( Z := \{ z \in \{0, 1\}^{\sum_{j=1}^L z_{j,t,o} \geq 1} \}. \) The approximate problem \( (4) \) involves \( n_c = T_nu \) continuous and \( n_b = LTO \) binary variables.

Given \( N \) independent and identically distributed samples \( X_{t,o} \) for \( t \in Z_{1:T} \) and \( o \in Z_{1:O}, \) we evaluate the constraints in problem \( (4) \) at the samples as follows:

\[
\begin{align*}
\min_{u \in U, z \in Z} & \quad J(u) \tag{5a} \\
\text{s.t.} & \quad (b), (c) \tag{5b} \\
& \quad A_j \left( X_{t,o}^{(i)} \right)^\top y_t + M \left( 1 - z_{j,t,o} \right) \geq b_j \left( X_{t,o}^{(i)} \right), \quad \forall j \in Z_{1:L}, t \in Z_{1:T}, o \in Z_{1:O}, i \in Z_{1:N}. \tag{5c}
\end{align*}
\]

We can apply the main result of [10].

**Lemma 1** ([10]): Given a risk level \( \epsilon \in (0, 1) \) and a confidence bound \( \beta \in (0, 1) \), let \( N \) satisfy

\[
2^{n_b} \sum_{i=0}^{n_c-1} \binom{N}{i} c_i (1-c)^{N-i} \leq \beta,
\]

where \( n_c = T_nu \) and \( n_b = LTO. \) Then, the optimal solution of \( (5) \) is feasible for \( (4) \) with probability at least \( 1 - \beta. \)

**C. Drawbacks of Scenario Problem (5)**

Although problem \( (5) \) is readily solvable by available mixed-integer program solvers, it has two drawbacks. First, the number of mixed-integer constraints \( (5c) \) is \( LTON. \) This means that \( (5) \) involves a large number of mixed-integer constraints because the lower bound of \( N \) that satisfies \( (4) \) is usually large. For example, in one of the simulations in Section IV \( N = 1706 \) and solving \( (5) \) takes about 4s, which indicates that this approach may not be suitable for real-time applications. Second, scenario problem \( (5) \) tends to be conservative, especially when \( X_{t,o} \) has a multimodal distribution. We illustrate this aspect in the following example.

**Example 1:** Consider a simple one-dimensional example where \( u = y \in \mathbb{R} \) and \( X_{t,o} \in \mathbb{R}. \) We let \( T = 1, O = 1 \) and omit the subscripts \( t, o \) in the random variable \( X_{t,o}. \) Suppose that the distribution of \( X \) takes the form of a mixture of two uniform distributions as depicted in Fig. \( 2(a) \), from which we draw \( N \) samples. Suppose that the maximum and minimum values of the samples are \( c \) and \( -c, \) respectively, for some positive constant \( c \in \mathbb{R}. \) Let \( \mathcal{O}(X) = (X - 0.1, X + 0.1) \subset \mathbb{R} \) and thus \( L = 2. \) The scenario problem \( (5) \) that minimizes \(|y|\) then becomes

\[
\begin{align*}
\min_{y \in \mathbb{R}, z \in Z} & \quad |y| \tag{7a} \\
\text{s.t.} & \quad y + M (1 - z_1) \geq X^{(i)} + 0.1, \quad i \in Z_{1:N}, \tag{7b} \\
& \quad y + M (1 - z_2) \geq -X^{(i)} + 0.1, i \in Z_{1:N}. \tag{7c}
\end{align*}
\]

where \( Z = \{(z_1, z_2) \in \{0, 1\}^2 : z_1 + z_2 \geq 1\}. \) By Lemma \( 1 \) if \( N \geq 190.6, \) the optimal solution of \( (7) \) guarantees \( (4c) \) with \( \epsilon = 0.05 \) with probability at least \( 1 - 0.01. \) Note that \( y = 0 \) avoids entering all the obstacle sets, i.e., \( y \not\in \mathcal{O}(X^{(i)}) \) for all \( i, \) as no sample lies in between \(-1 \) and \( 1. \) However, in problem \( (7), \) \( y = 0 \) is not a feasible solution because there is no associated feasible set of binary variables. Instead, the optimal cost of \( (7) \) is \( c + 0.1, \) which can be arbitrarily large.

Example \( 1 \) illustrates that scenario-based formulation \( (5) \) may result in a conservative solution due to an insufficient number of binary variables when uncertain parameters have a multimodal distribution. If we introduce an additional set of binary variables for each mode of the distribution, we can reduce the conservativeness. This is the intuitive idea behind our approach, which we will explain in the next section.

### III. Solution Approach

Our approach to solving problem \( (1) \) is to:

A. Separate samples into \( K \) clusters where \( K \) is the number of modes in the distribution.

B. Compute a set \( C^k \) that contains the obstacle sets in each cluster \( k. \)

C. Formulate a deterministic motion planning problem given \( \bigcup_{k=1}^K C^k. \)

In step A, our approach explicitly handles \( K \) modes of the distribution to reduce the conservativeness. We formulate step B as a computationally efficient linear program, and use the result to make the computation of the more challenging, mixed-integer motion planning problem in step C independent of \( N. \) In this section, we explain the three steps in detail.

### A. Clustering

There are several approaches to the clustering step. Given the number of modes \( K_o \) of the probability distribution for OV \( o, \) one approach is to use the \( k \)-means clustering method [16] to separate samples into \( K_o \) clusters. Note that modes are often known in many driving scenarios. For example, at intersections, vehicles can either go straight or turn, in which case clustering amounts to determining whether \( X_{t,o} \) corresponds to going straight or turning. Instead of running the clustering algorithm for each \( t \in Z_{1:T}, \) the clustering algorithm can be run for samples \( X_{T,o}^{(i)} \) \( i \in Z_{1:N} \) at the last prediction time step \( T. \) This is due to the observation that in almost all cases modes exhibit the greatest separation at the longest prediction time \( t. \) An alternative approach is to directly use the data provided from forecasting neural networks. In particular, such networks commonly use latent variables to encode driving networks. In particular, such networks commonly use latent variables to encode driving networks.
clusters: $X_t^{(i)}$ for $i \in I_o^k$, $k \in \mathbb{Z}_{1:K_o}$, where $I_o^k$ is the set of sample indexes in cluster $k$, $\bigcup_{k \in I_o^k} I_o^k = \emptyset$, $\bigcup_{k \in I_o^k} I_o^k = \mathbb{Z}_{1:N}$, and is the same for all time steps $t = 1, \ldots, T$.

Based on the above discussion, we assume that the clustering is a predetermined process.

Assumption 1: Clustering is done via a function $g : \mathbb{R}^n \rightarrow \mathbb{Z}_{1:K_o}$ that assigns each realization $X_{t,o} \in \mathbb{R}^{nx}$ to a corresponding cluster $g(X_{t,o}) \in \mathbb{Z}_{1:K_o}$.

We demonstrate the two ways mentioned above of defining $g$ in Section IV.

B. Polytopic Approximation

If a set $C_t,o$ satisfies $\bigcup_{i=1}^N O(X_t^{(i)}) \subseteq C_t,o$, then constraint $y_t \notin O(X_t^{(i)})$, $\forall i \in \mathbb{Z}_{1:N}$ is implied by $y_t \notin C_t,o$. For the tractable computation of the collision avoidance constraint $y_t \notin C_t,o$, we represent $C_t,o \subseteq \mathbb{R}^n$ by the union of polytopes $\bigcup_{k=1}^{K_o} C_t,o^k := \bigcup_{k=1}^{K_o} \{ y \in \mathbb{R}^n : A_t,o y \leq b_t,o \}$, where $C_t,o^k$ is a polytope that contains the obstacle sets in cluster $k$, $A_{t,o}^k \in \mathbb{R}^{L_{c,t,o} \times N}$, $b_{t,o}^k \in \mathbb{R}^{L_{c,t,o}}$, and $L_{c,t,o}$ is the number of halfspaces.

To efficiently obtain $C_t,o^k$, we predetermine $A_{t,o}^k$ and let $b_{t,o}^k$ be the only decision variable. For example, we set $A_{t,o}^k$ to be $A_{t}^k(X_{t,o}^k)$, the same matrix as in the obstacle set, evaluated at the element-wise average $\bar{X}_{t,o}^k := \frac{1}{|T|} \sum_{t \in T} X_{t,o}^k$. Moreover, we use the vertexes of $O(X_t^{(i)})$ to significantly speed up the computation. For example, the vertexes are easily computable when we model the obstacle set as a rectangle of the OV width and length, rotated with the OV yaw angle. Let $\mathcal{Y}_{t,o}^k$ be the set of the vertexes of $O(X_t^{(i)})$ for $i \in I_t^k$. Then, the overapproximation $C_t,o^k$ should contain all the vertexes in $\mathcal{Y}_{t,o}^k$ due to convexity. Following the previous problem computes $C_o^k = (C_{t,o}^1, \ldots, C_{t,o}^{K_o})$ for each cluster $k$ of OV $o$.

\[
\begin{align*}
\min_{b_{t,o}^k} & \, \sum_{t=1}^{T} \bar{J}(b_{t,o}^k) \\
\text{s.t.} & \quad \bar{A}_{t,o} v \leq b_{t,o}^k, \quad v \in \mathcal{Y}_{t,o}^k, t \in \{1:4,T\}.
\end{align*}
\]

(8a)

(8b)

We use $\bar{J}(b) = \sum_{t=1}^{T} b_t$ as a heuristic to yield a tight approximation because a larger $\bar{J}(b)$ then indicates a larger volume of polytope $\{ y : A \bar{y} \leq \bar{b} \}$ when $A$ is fixed. Problem (8) consists of $L_{c,T} \sum_o |\mathcal{Y}_{t,o}^k|$ constraints. Although the number of constraints increases with the number of samples $N$ since $|\mathcal{Y}_{t,o}^k| = O(LN)$, it can be efficiently solved by linear program solvers, or by applying an element-wise maximum operator to $A_{t,o}^k v$ in (8b) over all $v$ because with our choice of the cost function, the optimal solution is $b_{t,o}^k = \max_{v \in \mathcal{Y}_{t,o}^k} (\bar{A}_{t,o} v)$.

C. Mixed-integer Motion Planning Formulation

Using the overapproximation of the obstacle sets, we formulate a deterministic motion planning problem.

\[
\begin{align*}
\min_{u \in \mathcal{U}} & \, J(u) \\
\text{s.t.} & \quad (1b), (1c), \\
& \quad y_t \notin C_{t,o}^k, \quad k \in \mathbb{Z}_{1:K_o}, t \in \mathbb{Z}_{1:T}, o \in \mathbb{Z}_{1:O}.
\end{align*}
\]

(9a)

(9b)

(9c)

Problem (9) can be written as a mixed-integer program. It contains $L_{c,T} \sum_o K_o$ binary variables and $L_{c,T} \sum_o K_o$ mixed-integer constraints (9c). Compared with scenario problem (5), which involves LTO binary variables and LTON mixed-integer constraints (53), problem (9) requires more binary decision variables because $L$ is similar to $L_{c}$ and $K_o > 1$ if the distribution is multimodal, but it significantly reduces the number of mixed-integer constraints as the number of clusters is much smaller than the number of samples, i.e., $\sum_o K_o \ll ON$.

Remark 1: Recall that we use a linear time-varying model for the EV in (1b) and (1c). To handle a nonlinear dynamic model, we can linearize it at each time step to obtain a linear time-varying model, and use the same motion planning formulation as (9). In fact, this stepwise linearization is a common approach in vehicle control to handle nonlinear dynamics [17]. For OVs, any dynamic model can be used because we only use samples of their predicted trajectories to construct the polytopic constraints (9c).

We now present the main result that generalizes a feasible solution of (9) to that of chance-constrained problem (1).

Theorem 1: Suppose that Assumption 1 holds true. Given a risk level $\epsilon \in (0,1)$ and a confidence bound $\beta \in (0,1)$, let $c_{k}^o \in (0,1)$ and $b_{k}^o \in (0,1)$ be chosen such that

\[
\sum_{o=1}^{O} \sum_{k=1}^{K_o} c_{k}^o = \epsilon, \quad \sum_{o=1}^{O} \sum_{k=1}^{K_o} b_{k}^o = \beta.
\]

(10)

Let the number of samples $N_{k}^o$ satisfy (9) with $c_{k}^o, b_{k}^o, n_c = L_{c,T}$ and $n_o = 0$. Then, a feasible solution of (9) is a feasible solution of (1) with probability at least $1 - \beta$.

Proof: Let $y^* := (y_t^*, \ldots, y_T^*)$ denote the output corresponding to (1b) and (1c) from a feasible solution of (9), and $C_{k}^o$ denote the optimal overapproximation of (9). For each $t$ and $o$, let us partition the set of samples into $S_{t,o}^k, o \in \mathbb{Z}_{1:O}$, $k \in \mathbb{Z}_{1:K_o}$, where $S_{t,o}^k := \{ X_{t,o} \in \mathbb{R}^{nx} : g(X_{t,o}) = k \}$. Such a partition is given by Assumption 1. Let $S_{t,o}^k := S_{t,o}^k \cap \mathbb{X}_{t,o}$.

To prove the theorem, we want to show $Pr^N(y^* \leq \epsilon) \geq 1 - \beta$ where $N = \max(N_{k}^o)$ and $V(y^*) := Pr(\mathcal{X} : \exists t, o \text{ such that } y_t^* \in O(X_{t,o}) \text{ here), Pr}^N$ is the probability measure of $N$ samples $X^O(\bar{U})$. To do this, we use $V(C_{k}^o) := Pr(X_o \in S_{t,o}^k : \forall t^* \text{ such that } O(X_{t,o}) \subseteq C_{t,o}^k) \text{ and bound } V(y^*) \text{ in terms of } V(C_{k}^o)$ so that eventually we can bound $Pr^N(y^* \leq \epsilon)$ in terms of $Pr^N(V(C_{k}^o) \leq \epsilon^k)$. Since $y^*$ is a feasible solution of (9), we have $y_t^* \notin C_{k}^o$ for any $t$, $o$, and $k$. Thus, $O(X_{t,o}) \subseteq C_{t,o}^k$ implies $y_t^* \notin O(X_{t,o})$, which leads to $1 - V(y^*) = Pr(X : A_{t,o} y_t^* \notin O(X_{t,o})) \geq Pr(A_{t,o} O(X_{t,o}) \subseteq C_{t,o}^k) = 1 - \sum_o Pr(X_o \notin O(X_{t,o}) \subseteq C_{t,o}^k) = 1 - \sum_o \sum_{k \neq k} Pr(X_o \in S_{t,o}^k : \forall t \text{ such that } O(X_{t,o}) \subseteq C_{t,o}^k) = 1 - \sum_o \sum_{k \neq k} V(C_{t,o}^k).

The second inequality is by Boole’s inequality, i.e.,

$Pr(V(A_{t,o} y_t^* \notin O(X_{t,o}) \subseteq C_{t,o}^k)) \leq 1 - \sum_o \sum_{k \neq k} V(C_{t,o}^k).$

Since $N_{k}^o$ satisfies (9), by Lemma 1 $Pr^N(V(C_{k}^o) \leq \epsilon^k) \geq 1 - \beta$. This allows us to complete the proof because $Pr^N(V(y^*) \leq \epsilon) \geq Pr^N(V(C_{k}^o) \leq \epsilon^k) \geq Pr^N(\mathcal{A}_{o,k} V(C_{k}^o) \leq \epsilon^k) \geq 1 - \sum_o \sum_{k \neq k} Pr^N(V(C_{k}^o) \leq \epsilon^k) \geq 1 - \beta$. ■
Theorem 1 illustrates that if we have \( N_o^k \) samples to compute \( C_o^k \) for each OV \( o \) and cluster \( k \), then any feasible solution of our approach (i.e., (8) and (9)) is a feasible solution to (1) with probability at least 1 − \( \beta \).

In Theorem 1 \( e_o^k \) and \( \beta_o^k \) that satisfies (10) can be chosen based on problem-specific knowledge. For example, in the case study in Section IV-B, the probability \( p(g(X_{t,o}) = k) \) is known. In that case we can set \( e_o^k \) for each OV to be proportional to the inverse of \( p(g(X_{t,o}) = k) \) so that a less likely cluster is assigned a higher risk and thus requires fewer samples. Another choice is to allocate \( \epsilon \) and \( \beta \) uniformly over clusters: \( e_o^k = \epsilon/\sum_{k=1}^{K_o} K_o \) and \( \beta_o^k = \beta/\sum_{k=1}^{K_o} K_o \).

Example 2: Recall Example 1. Now, we separate the samples into two clusters and compute \( C_1^2 \) and \( C_2^2 \) by (8) given predetermined vectors \( A_1^2 = A_2^2 = (1, -1)^T \). Then (9) becomes the problem that minimizes \( |y| \) subject to \( y \notin C_1^2 \) and \( y \notin C_2^2 \). By Theorem 1 with the uniform allocation of \( \epsilon \) and \( \beta \), if \( N_1^2 = N_2^2 \geq 400.6 \), any feasible solution of the problem ensures \( \Pr(y \notin O(X)) \geq 1 - 0.05 \) with probability at least 1 − 0.01. The optimal cost is 0 because \( y = 0 \) is feasible, as illustrated in Fig. 2(b). Compared with Example 1, our approach yields a less conservative solution.

In the next section, we will show that even for a higher dimensional practical example, our approach is less conservative and computationally more efficient than the standard scenario approach without preprocessing.

IV. SIMULATION RESULT

In this section, we compare the performance of our approach to that of the basic scenario approach via simulations. We model the EV dynamics as double integrators with speed and input constraints as in the case studies of [4]. In our first case study, we impose multimodal uncertainties on the acceleration of the OV to generate trajectory predictions. In the second case study, we integrate our approach with the state-of-the-art forecasting neural network Trajectron++ [6].

A. Case study 1: Lane change

In Fig. 3 the EV intends to change lane, and the OV is predicted to decelerate or accelerate. The OV’s predicted potential obstacle sets are shown in red, and the EV’s position resulting from our approach and the basic scenario approach is shown in blue and magenta, respectively. The objective here is to progress furthest to the right while achieving the bottom lane. With \( \epsilon = 0.05, \beta = 10^{-3}, L_c = 4, T = 10, O = 1, K_1 = 2 \) and uniform allocation of \( \epsilon \) and \( \beta \), we use \( N_1^2 = N_2^2 = 2964 \) trajectories (in total \( N = 5928 \)) based on Theorem 1 for our scheme. As reported in Table I the basic scenario approach requires fewer samples, but our approach is more than 20 times faster because the number of samples \( N \) affects only the computation time of the linear program (8), not that of the mixed-integer program (9). The reported computation time does not include sample generation time. In our approach, clustering, (8), and (9) take 0.06 s, 0.02 s, and 0.1 s, respectively (summing to 0.18 s). Also, our approach is less conservative in the sense that it yields a better objective cost. Indeed, Fig. 3 shows that the EV slows down to change lanes in the basic scenario approach, whereas in our approach it neatly fits itself between the two possible OV behaviors. For the optimal output \( y^* \) of our approach, we checked whether \( y_t^* \notin O(X_{t,o}^{(1)}) \) for all \( t \in Z_{1:T} \) for 10^5 newly predicted trajectories. The empirical violation—the fraction of sampled cases which demonstrate violations—is higher in our approach because it enables less conservative motions, but is still within the specified risk level of \( \epsilon = 0.05 \). In some cases (e.g., when the EV initial velocity is large), our approach fails in finding a lane-changing trajectory, but still guarantees safety at the risk level of \( \epsilon \) with high confidence.

B. Case study 2: Intersection on the nuScenes dataset

We train the forecasting neural network Trajectron++ [6] on the nuScenes dataset [12] using Compute Canada and

\( \text{https://www.computecanada.ca/} \)
We have presented a motion planning approach that ensures probabilistic safety in the presence of multimodal uncertainties. We have formulated the motion planning problem as a chance-constrained problem and presented an efficient, sampling-based approach to obtain a high-confidence solution. In particular, we have characterized the number of samples required to guarantee probabilistic safety with high confidence. We have validated on a real-world dataset that our approach is less conservative and computationally more efficient than a conventional scenario approach. Since our approach works for any multimodal distribution from which we can sample, it could be integrated with, for example, prediction models based on Gaussian processes or Markov Decision Processes. We are currently implementing a receding horizon version of our approach to achieve closed-loop control in dynamic environments.

V. Conclusion

We have presented a motion planning approach that ensures probabilistic safety in the presence of multimodal uncertainties. We have formulated the motion planning problem as a chance-constrained problem and presented an efficient, sampling-based approach to obtain a high-confidence solution. In particular, we have characterized the number of samples required to guarantee probabilistic safety with high confidence. We have validated on a real-world dataset that our approach is less conservative and computationally more efficient than a conventional scenario approach. Since our approach works for any multimodal distribution from which we can sample, it could be integrated with, for example, prediction models based on Gaussian processes or Markov Decision Processes. We are currently implementing a receding horizon version of our approach to achieve closed-loop control in dynamic environments.

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