Load-Aware Modeling and Analysis of Heterogeneous Cellular Networks

Harpreet S. Dhillon, Student Member, IEEE, Radha Krishna Ganti, Member, IEEE, and Jeffrey G. Andrews, Fellow, IEEE

Abstract—Random spatial models are attractive for modeling heterogeneous cellular networks (HCNs) due to their realism, tractability, and scalability. A major limitation of such models to date in the context of HCNs is the neglect of network traffic and load: all base stations (BSs) have typically been assumed to always be transmitting. Small cells in particular will have a lighter load than macrocells, and so their contribution to the network interference may be significantly overstated in a fully loaded model. This paper incorporates a flexible notion of BS load by introducing a new idea of conditionally thinning the interference field. For a K-tier HCN where BSs across tiers differ in terms of transmit power, supported data rate, deployment density, and now load, we derive the coverage probability for a typical mobile, which connects to the strongest BS signal. Conditioned on this connection, the interfering BSs of the kth tier are assumed to transmit independently with probability pk, which models the load. Assuming – reasonably – that smaller cells are more lightly loaded than macrocells, the analysis shows that adding such access points to the network always increases the coverage probability. We also observe that fully loaded models are quite pessimistic in terms of coverage.

Index Terms—Heterogeneous cellular networks, load-aware model, Poisson point process, stochastic geometry, HetNet performance analysis.

I. INTRODUCTION

Advances in hardware and the increasing popularity of smartphones and tablets have led to a paradigm shift in the way cellular networks are accessed and consequently the way they are deployed. In terms of network access, focus has shifted from voice-oriented applications towards data-hungry applications such as live video streaming and symmetric video calls [3]. Macrocell based conventional cellular networks were primarily designed to provide coverage and are clearly not capable of accommodating this huge change in the usage trends [3]. One of the most promising ways to handle this data deluge is to increase the density of the BSs, thereby reducing the frequency reuse distance and hence improving network capacity [4]. For example, a typical 3G or 4G cellular network already has operator-managed picocells deployed at the hot spots and cell edges [5]; distributed antennas deployed to eliminate coverage dead-zones [6]; and low-power user-deployed femtocells [3]; along with the existing high-power tower-mounted traditional macrocell BSs that guarantee universal coverage.

A. Related Work and Motivation

The rapidly increasing heterogeneity and randomness in cellular networks threaten classical models, such as the hexagonal grid [7] or Wyner model [8], with obsolescence. Clearly, a more sensible way to model a HCN is with a random spatial model, where the BS locations form a realization of some random spatial point process [9]–[12]. Such a model captures the inevitable uncertainty in their locations, and tools from stochastic geometry [13] and point process theory [14] can be deployed to assist in analysis [15].

For example, in an HCN, macrocells would usually follow a somewhat sparse point process and have high transmit power, whereas pico and femtocells are drawn from successively denser (more BSs/area) processes, and have lower transmit power. In such a network, a mobile user could simply connect to the strongest base station signal, with the rest of the transmitting BSs being interferers. This model was introduced in [16], [17] and extended in [18]–[20], and is surprisingly tractable: under fairly benign assumptions, the coverage probability could be derived in closed-form, which is not possible even for 1-tier networks in the hexagonal grid model. The model further was shown to generally agree in several important ways with more sophisticated industry (e.g. 3GPP) simulations [21] and even early field deployments of HCNs [22].

Despite this encouraging progress, these models lack in at least one important aspect, which is their neglect of network traffic and load. Rather, the work to date in this direction has assumed that all the BSs transmit concurrently all the time, which translates to a fully loaded (or full buffer) scenario resulting in pessimistic estimates of coverage and average rate. Although this might be justified for macrocells in peak traffic hours, this is not applicable for smaller cells whose smaller coverage areas will naturally accommodate fewer users, even if considerable biasing towards the small cells is introduced. Therefore, the main goal of this paper is to incorporate a notion of BS load. Those familiar with random spatial models will recognize that a simple independent thinning of the point processes will not capture the load since it may also turn off the serving BS, which is not allowed if the analysis is performed for a typical active user. On the other
hand, incorporating more sophisticated queueing models in the present multi cell scenario will render the analysis intractable due to the interference induced coupling in the service rates of various BSs [23], [24]. Moreover, this line of thought is not in the scope of the current paper since we do not focus on the flow level performance evaluation. The readers interested in flow level models can refer to [25], [26]. With our main focus on the downlink coverage evaluation, we propose a middle way whereby we conditionally thin the interference field predicated on a connection to a typical active user, and we are able to maintain acceptable tractability with a realistic model of BS loading.

B. Contributions and Outcomes

The main contributions of this paper are as follows:

1) Tractable Load Model for K-Tier HCNs: In Section II we incorporate a notion of BS load in a general K-tier random spatial model for HCNs. For an HCN where BSs across tiers differ in terms of their transmit power, supported data rate and deployment density, we assume that a typical mobile connects to the strongest BS in terms of received power and conditioned on this connection, the i\textsuperscript{th} tier interfering BSs transmit independently with a probability p\textsubscript{i}, which models the load. These BS activity factors \{p\textsubscript{i}\} may vary significantly across the tiers due to different coverage areas of each tier.

2) Coverage Probability: We derive exact expressions for the coverage probability of a typical mobile user in both open and closed access HCNs. Since these expressions involve an infinite summation, we also derive a set of upper and lower bounds that can be made arbitrarily tight with a finite number of terms. These bounds also give insights into the number of terms of the infinite summation required to approximate the coverage probability such that the approximation error is within some predefined limit.

3) Design Insights: This paper provides some potentially useful design insights for HCNs. First, we study the effect of proposed “conditional thinning” on the coverage footprints of various tiers and show that this can be understood in two equivalent ways: i) thinning of interference, and ii) biasing of the typical mobile towards its serving BS. While the former is a direct result of thinning, the latter is an indirect consequence of the expansion of the coverage regions in the thinned interference field.

Second, our analysis sheds light into the effect of adding new tiers to already existing HCNs. In particular, we derive an exact condition under which the addition of a new tier to a general K-tier HCN will increase the overall coverage probability. A relevant special case is the addition of small cells to existing macrocell networks, where we show that in the interference limited regime the overall open access coverage probability increases if the load on small cells is smaller than that of macrocells, which is a typical operating scenario because of the smaller loads handled by small cells. This is a strong rebuttal to the viewpoint that unplanned infrastructure might bring down a cellular network due to increased interference.

Third, we show that the coverage probability for a general K-tier interference-limited open-access network is invariant to changes in the power and deployment density when all the classes of BSs have same loads and target SINR\textsuperscript{s}. Furthermore, this coverage probability is also the same as that of a single tier network with the same target SINR and the same BS activity factor.

II. SYSTEM MODEL

We model a downlink heterogeneous cellular network with K classes (or tiers) of BSs. For notational simplicity, we denote the set \{1, 2, \ldots, K\} by K. BSs of the i\textsuperscript{th} class transmit with power \(P_i\), have a target SINR of \(\beta_i\) and are assumed to form a realization of an independent homogeneous Poisson Point Process (PPP) \(\Phi_i\) with density \(\lambda_i\). Such a model seems sensible for user deployed BSs such as femtocells but is dubious for the centrally planned tiers such as macrocells. Nevertheless, the difference is not as large as expected and PPP assumption for macrocells is shown to be about as accurate as the grid model when compared to an actual 4G network in [11]. More recently, [27] has validated the PPP assumption for certain cities using tools from spatial statistics. Furthermore, this model is likely even more sensible for K-tier HCNs due to the increased uncertainty in the deployment of lower tiers (smaller cells). We will comment more on the accuracy of this assumption in the light of the proposed load model in the Numerical Results Section.

Without loss of generality, we perform analysis on a typical mobile user located at origin, which is made possible by Slivnyak’s Theorem [13]. For cell association, we consider the max-SINR connectivity model, where a mobile user connects to the BS that provides highest downlink SINR. It should be noted that this model is the same as the max-power connectivity model where a mobile connects to the BS that provides highest downlink power. Since HCNs are typically interference-limited [28], we ignore thermal noise for notational simplicity. However, as would be evident from the analysis, this assumption can be relaxed without much extra work. To model the wireless channel, we consider a standard distance based path loss with exponent \(\alpha > 2\) along with Rayleigh fading. Hence the received power at a typical mobile from a BS located at point \(x \in \Phi_i\) can be expressed as \(P_i h_x \|x\|^{-\alpha}\), where \(h_x \sim \exp(1)\) and \(\|x\|^{-\alpha}\) is the distance based path loss. General fading distributions, e.g., log-normal shadowing, can be incorporated using techniques developed in [29] at the cost of tractability. Assuming \(Z_k\) to be the set of \(k\)-tier interfering BSs (possibly thinned version of \(\Phi_i\)), the downlink SINR at the typical mobile user when it connects to the BS located at point \(y \in \Phi_i\) is:

\[
\text{SINR}(y) = \frac{P_i h_y \|y\|^{-\alpha}}{\sum_{k=1}^{K} \sum_{x \in \Phi_k} P_k h_x \|x\|^{-\alpha}}. \tag{1}
\]

A. Modeling Base-Station Load

In this K-tier random spatial model, we now incorporate network “load” perceived by each BS as the likelihood of its transmission at a randomly chosen time instant. This can also be visualized as the BS activity factor, formally defined as the fraction of time for which a BS transmits.
Relationship of BS activity factor with number of active users: A BS is inactive in a particular resource block, e.g., time-frequency resource block in LTE [30], if there is no active user scheduled. This can be due to an over provisioned system or a momentary lull in traffic due to the bursty nature of data access. Clearly, this model characterizes the load on each BS in terms of the total number of active users served by that BS at a random time instant. In the context of Orthogonal Frequency Division Multiple Access (OFDMA) if a particular BS experiences high load, it will utilize more frequency time resources and hence the probability that a user is scheduled in a particular frequency time block increases. Therefore, the load perceived by a BS is directly related to the likelihood that an arbitrary resource block is utilized and hence is related to the BS activity in that particular block.

Temporal and spatial correlation in BS activity factors: In general, there is both temporal and spatial correlation in the activity factors of different BSs. Temporal correlation is induced across neighboring BSs by the mobility of users, i.e., if a user is associated to a particular BS, the likelihood of neighboring BSs transmitting at a future time instant is slightly higher. Spatial correlation is induced by interference and traffic/load patterns [23], [24]. To understand this, consider two neighboring BSs. When the first BS transmits, it increases net interference experienced by the second BS and hence reduces its data rate. As a result, the second BS now takes longer to transmit same amount of data than it would have taken if the first BS was not transmitting. Therefore, the activity factors of these two BSs are positively correlated. However, modeling the exact nature of these correlations is beyond the scope of the current paper and we assume the BS activity factors to be independent. Although the spatio-temporal correlations haven’t yet been modeled for this exact problem, it is worth noting that they have been handled in some related setups, e.g., the effect of spatio-temporal correlations of interference on coverage is discussed in [31], [32].

B. Proposed Load Model and Mathematical Preliminaries

We assume that a typical mobile connects to the strongest BS in terms of received power and conditioned on this connection, the interferer belonging to the \( i \)-th tier transmits independently with a probability \( p_i \) and is idle with a probability \( 1 - p_i \). This conditioning makes it harder to analyze this system model since we do not have \textit{a priori} knowledge about the serving BS and hence it is not possible to isolate the interference field. To overcome this, we partition each tier \( \Psi_m \) independently into two sets of BSs \( \Psi_m \) and \( \Delta_m \), where \( \Psi_m \) and \( \Delta_m \) are both independent PPPs with densities \( p_m \lambda_m \) and \( (1 - p_m) \lambda_m \). The set \( \Psi_m \) represents the set of active BSs of tier \( m \) with the possibility of one of them being a serving BS, and \( \Delta_m \) represents the set of idle BSs of tier \( m \) with an exception that it could also contain the serving BS since partitioning was done independently. The advantage of this partitioning is that the interferers are confined to the set \( \Psi = \bigcup_{m \in K} \Psi_m \). For ease of notation, we define the maximum signal strength from a set of nodes \( \mathcal{A} \) as

\[
M(\mathcal{A}) = \sup_{x \in \mathcal{A}} P_A h_x ||x||^{-\alpha},
\]  

and the total received power at the origin from the set of active BSs as:

\[
I = \sum_{i=1}^{K} \sum_{x \in \Psi_i} P_i h_x ||x||^{-\alpha},
\]

which denotes the net interference power if \( \Psi \) does not include the serving BS and the interference plus signal power if it includes the serving BS. From the definition of \( M(\Psi_i) \) and \( I \), it is easy to see that \( 1 \left( \frac{M(\Psi)}{I - M(\Psi)} < \beta_i \right) = 1 \) only if no active BS in the set \( \Psi_i \) can connect to the mobile. Similarly, \( 1 \left( \frac{M(\Delta)}{I} < \beta_i \right) = 1 \) only if no BS in the set \( \Delta_i \) is able to connect to the mobile. The second event is defined to cover the possibility that a serving BS may lie in the set \( \Delta_i \). Using these two events, we will now define the coverage probability of a typical mobile at the origin. We note that a mobile will be in outage (not in coverage) if none of the BSs in the whole network provides SIR that is greater than the corresponding target for that tier.

**Definition 1** (Coverage Probability). Coverage probability, \( P_c \), can be formally defined as:

\[
P_c = 1 - \mathbb{E} \left[ \prod_{i \in K} 1 \left( \frac{M(\Psi_i)}{I - M(\Psi)} < \beta_i \right) \right] \left[ \frac{M(\Delta_i)}{I} < \beta_i \right].
\]

For this definition, we implicitly assumed an open access network where a mobile user is allowed to connect to any BS in the network without any restrictions. Another possible access strategy is closed access or closed subscriber group strategy in which a mobile is allowed to connect to only a subset \( \mathcal{B} \subseteq \mathcal{K} \) of all the tiers. Coverage probability for closed access is also given by (4) with the only difference that the product is over the set \( \mathcal{B} \) instead of \( \mathcal{K} \).

For tractability, we assume that the target SIR thresholds \( \beta_i \) are greater than 0 dB, i.e., \( \beta_i > 1, \forall i \). This is in fact the case for a large fraction of mobile users and only a few edge users might violate this assumption. Moreover, in the Numerical Results Section we show that the results derived under this weaker assumption hold down until around \(-2\) dB which covers a large fraction of cell edge users as well. This assumption has also been validated earlier for the fully loaded \( K \)-tier HCN in [17]. The reason why this assumption is helpful is because it ensures that at most one BS in the active set \( \Psi \) meets the target SIR requirements for a typical mobile user. Refer to [17] for a detailed discussion on this assumption and its application in coverage analysis of a fully loaded \( K \)-tier HCN.

C. Coverage Regions

Before going into the detailed analysis of coverage probability, it will be useful to understand the effect of the proposed load model on the coverage footprints of various BSs. Consider a realization of a three tier HCN in Fig. We first plot the coverage regions assuming a fully loaded network by tessellating the space according to max-SIR connectivity model in the left figure. Clearly, this plot does not resemble a classical Voronoi tessellation due to the differences in the
transmit powers of BSs across tiers. Moreover, it should be noted that the “cell edges” are not as sharp in reality due to fading and shadowing, which are averaged out for these illustrative plots. The effect of incorporating the proposed load model on coverage footprints can now be understood in two equivalent ways: i) thinning of the interference field conditional on the connection of a typical mobile to its serving BS, where the original coverage regions corresponding to the inactive BSs are removed to highlight conditional thinning (middle figure), ii) biasing of a typical mobile towards its serving BS relative to the new cell edge defined by the set of active BSs (right figure). While the former is a direct result of conditional thinning, the latter is an indirect consequence of the expansion of coverage regions in the thinned interference field.

III. COVERAGE PROBABILITY

This is the main technical section of this paper where we derive the probability that a typical mobile is in coverage under the system model introduced in the last section. We first derive coverage probability for an open access network, from which the results for closed-access immediately follow.

A. Exact Expression for Coverage Probability

We start by stating the Laplace transform of $I$, i.e., $\mathcal{L}_I(s) = \mathbb{E}[\exp(-sI)]$, in Lemma 1, which will be useful in the derivation of coverage probability. The proof is given in [17].

**Lemma 1.** The Laplace transform of $I$ can be expressed as:

$$\mathcal{L}_I(s) = \exp \left( -s^\frac{2}{\alpha} C(\alpha) \sum_{i=1}^{K} p_i \lambda_i P_i^\frac{2}{\alpha} \right),$$

where $C(\alpha)$ is given by:

$$C(\alpha) = \frac{2\pi^2 \csc \left( \frac{2\pi}{\alpha} \right)}{\alpha}.$$  

The following Lemma deals with fractional moments of interference and is the main technical result required to evaluate the coverage probability for this model.

**Lemma 2.** Let $\Psi_i$ denote the set of active transmitters of tier $i$ and $\delta_i = \beta_i/(1+\beta_i)$. Let $I$ denote the total received power from the BSs in the set $\Psi$ and for notational simplicity define $T = 1 \left( \max_{i \in \mathcal{K}} \frac{\lambda_i(\Psi)}{\alpha} < I \right) I^{-2/\alpha}$. Then

$$\mathbb{E}[T^m] = \frac{m!g(m)}{(-A)^m},$$

where

$$g(m) = \left( -\frac{A}{\eta} \right)^m \left\{ \frac{1}{\Gamma(1+\frac{2m}{\alpha})} - \frac{B}{\Gamma(1+\frac{(m+1)\alpha}{\alpha})} \right\},$$

and

$$A = \pi \Gamma \left( 1 + \frac{2}{\alpha} \right) \sum_{i \in \mathcal{K}} (1 - p_i) \lambda_i P_i^\frac{2}{\alpha} \beta_i^{-\frac{2}{\alpha}},$$

$$B = \sum_{i \in \mathcal{K}} \frac{\lambda_i p_i P_i^\frac{2}{\alpha} \beta_i^{-\frac{2}{\alpha}}}{(1 + \beta_i) \frac{2m}{\alpha} + 1 + \frac{(m+1)\alpha}{\alpha}},$$

$$\eta = C(\alpha) \sum_{i=1}^{K} p_i \lambda_i P_i^\frac{2}{\alpha}.$$  

The hypergeometric function is denoted by $2F_1(a, b, c, z) = \frac{\Gamma(c)}{\Gamma(b)\Gamma(c-b)} \int_{0}^{1} (1-t)^{b-1}(1-t)^{-b-1} t^{a-c} \frac{1}{(1-t)^z} dt$.  

*Proof: See Appendix A.*

Using these Lemmas, we now derive the main coverage probability result.

**Theorem 1** (Open Access). The downlink coverage probability for a typical mobile user in a $K$-tier open access network assuming $\beta_i > 1, \forall i$, is

$$P_c = \frac{\pi}{C(\alpha)} \sum_{i \in \mathcal{K}} p_i \lambda_i P_i^\frac{2}{\alpha} \beta_i^{-2/\alpha} - \sum_{m=1}^{\infty} g(m),$$

*Proof: The coverage probability is given by (4). Since the point processes $\Delta_i$ and the corresponding fading random
variables are independent, conditioning on the common interference, we can move the expectation inside the product. Hence
\[
1 - P_c = \mathbb{E} \left[ \prod_{i=1}^{K} 1 \left( \frac{M(\Psi_i)}{I - M(\Psi)} < \beta_i \right) \mathbb{E} \left[ 1 \left( M(\Delta_i) < \beta_i I \right) \right] \right],
\]  
where the inner expectation is with respect to the inactive transmitter sets. We first simplify this inner expectation as follows:
\[
\mathbb{E} \left[ 1 \left( M(\Delta_i) < \beta_i I \right) \right] = \mathbb{E} \left[ \prod_{x \in \Delta_i} 1 \left( P_i h_k ||x||^{-\alpha} < \beta_i I \right) \right] \tag{13}
\]
\[
= \mathbb{E} \left[ \prod_{x \in \Delta_i} \left( 1 - \exp \left( -\beta_i P_i^{-1} I ||x||^\alpha \right) \right) \right] \tag{14}
\]
\[
= \exp \left( -\left(1 - p_i \right) \lambda_i \int_{x \in \mathbb{R}^2} \exp \left( -\beta_i P_i^{-1} I ||x||^\alpha \right) dx \right) \tag{15}
\]
\[
= \exp \left( -(1 - p_i) \lambda_i \beta_i^{-\frac{\alpha}{2}} I^{-\frac{\alpha}{2}} P_i^{\frac{\alpha}{2}} \pi \Gamma \left( 1 + \frac{2}{\alpha} \right) \right), \tag{16}
\]
where (a) follows from the fact that fading is Rayleigh distributed, i.e., \( h \sim \exp(1) \), (b) follows from the probability generating functional (PGFL) of PPP \( 13 \) and (c) follows from some algebraic manipulations to reduce the integral to a Gamma function. Now recalling the expression of \( A \) given by \( 8 \), we can write:
\[
1 - P_c = \mathbb{E} \left[ 1 \left( \max_{i \in \mathcal{K}} \frac{M(\Psi_i)}{\delta_i} < I \right) \exp(-AI^{-2/\alpha}) \right]. \tag{17}
\]
Using the Taylor series expansion of \( \exp(-x) \), exchanging the infinite summation and expectation\( 4 \):
\[
1 - P_c = \sum_{m=0}^{\infty} \frac{(-A)^m}{m!} \mathbb{E} \left[ 1 \left( \max_{i \in \mathcal{K}} \frac{M(\Psi_i)}{\delta_i} < I \right) I^{-2m/\alpha} \right]. \tag{18}
\]
The summation can be split as:
\[
1 - P_c = P \left( \max_{i \in \mathcal{K}} \frac{M(\Psi_i)}{\delta_i} < I \right) + \sum_{m=1}^{\infty} \frac{(-A)^m}{m!} \mathbb{E} \left[ T^m \right]. \tag{19}
\]
The term \( 1 - P \left( \max_{i \in \mathcal{K}} \frac{M(\Psi_i)}{\delta_i} < I \right) \) is the coverage probability in a fully loaded heterogeneous network where the \( m \)-th tier density is \( p_m \lambda_m \). This is derived in \( 17 \) and is given by:
\[
1 - P \left( \max_{i} \frac{M(\Psi_i)}{\delta_i} < I \right) = \frac{\pi}{C(\alpha)} \sum_{i=1}^{K} p_i \lambda_i I_2^{2/\alpha} \beta_i^{-2/\alpha} \tag{19}
\]
Using Lemma 2 to evaluate \( \mathbb{E} \left[ T^m \right] \), we obtain the result. \( \blacksquare \)

We note that the expression of coverage probability involves infinite summation over the sequence \( g(m) \). Therefore, we first show that the infinite summation converges by showing that \( |g(m)| \to 0 \) as \( m \to \infty \). Observe that:
\[
|g(m)| \leq \left( \frac{A}{\eta} \right)^m \frac{1}{\Gamma(1 + \frac{2m}{\alpha})} \leq \left( \frac{A}{\eta} \right)^m \frac{1}{\Gamma \left( 1 + \frac{2m}{\alpha} \right)} \tag{19}
\]
\[1^{\text{The average of the series is absolutely convergent.}}\]
B. Special Cases of Interest

We now use the results derived in this section to study some special cases and compare the system performance with already known results for fully loaded system. First, we note that for a fully loaded system, the value of $A = 0$ and hence $g(m) = g_{c}(m) = 0, \forall m$. Therefore, the coverage probability in this case can be expressed as the following Corollary of Theorems 1 and 2.

**Corollary 1 (Fully Loaded).** For a fully loaded system, i.e., $p_i = 1, \forall i$, the coverage probability in open access is given by:

$$P_c = \frac{\pi \beta^{-2/\alpha}}{C(\alpha)} \sum_{i \in K} \frac{\lambda_i P_i^{2/\alpha}}{\beta_i^{-2/\alpha}},$$

(23)

which is the same as Corollary 1 in [17]. The coverage probability in closed access is also given by (23) with the only difference that the summation over the set $K$ is now over set $B$.

For a single tier open access network, the coverage probability derived in Theorem 1 can be simplified and is expressed as the following Corollary.

**Corollary 2 (Single Tier).** The coverage probability for the single tier open access network with BS activity factor $p$ is

$$P_c = \frac{\pi \beta^{-2/\alpha}}{C(\alpha)} - \sum_{m=1}^{\infty} g(m),$$

(24)

where the terms $\frac{A}{\eta}$ and $\frac{B}{\eta}$ appearing in the expression of $g(m)$ given by (7).

$$A = \frac{\pi \Gamma(1 + \frac{2}{\alpha}) (1 - p)}{C(\alpha) \beta^{2/\alpha}}$$

(25)

$$B = \frac{2 F_1(1, \frac{2m}{\alpha}, 1 + \frac{(m+1)}{\alpha}, 1)}{C(\alpha) \beta^{2/\alpha} (1 + \beta)^{\frac{2m}{\alpha}}}.$$  

(26)

**Remark 1 (Scale invariance of a single tier network).** From Corollary 2 we note that for any BS activity factor $p$, the coverage probability in a single tier open access network is independent of the BS density $\lambda$ and transmit power $P$. This is henceforth referred to as “scale-invariance” of cellular networks to changes in the BS density and their transmit powers.

Remark 1 is a generalization of a similar result derived for fully loaded networks in [17], which can easily be seen from Corollary 1. In addition to single tier networks, it was also observed in [17] that the general fully loaded open access multi tier networks also exhibit scale invariance if the target SIRs for all the tiers are the same. This can also be easily deduced from Corollary 1. Motivated by this observation, we study the coverage probability for our proposed load model in open-access multi tier networks under the assumption that the target SIR is the same for all tiers in the next Corollary.

**Corollary 3 (Coverage Probability: $K$-Tier with same $\beta$).** The coverage probability for a $K$-tier open access network under the proposed load model assuming target SIRs to be the same ($= \beta$) for all the tiers is given by (24), with the difference that the term $\frac{A}{\eta}$ appearing in the expression of $g(m)$ given by (7) is:

$$A = \frac{\pi \Gamma(1 + \frac{2}{\alpha})}{C(\alpha) \beta^{2/\alpha}} \sum_{l=1}^{K} (1 - p_l) \lambda_i P_i^{2/\alpha} - \frac{B}{\eta},$$

(27)

and $\frac{B}{\eta}$ appearing in (7) is given by (26).

**Remark 2 (Scale invariance of $K$-tier HCNs with same $\beta$).** From Corollary 3 we note that the coverage probability for $K$-tier HCNs is not scale invariant in general, even when target SIRs of all the tiers are the same. However, the invariance property does hold when the BS activity factors of all the tiers are the same. Interestingly, the coverage probability in this case is same as that of a single tier network given by Corollary 2.

To understand this remark, we consider the following simple example.

**Example 1 (Scale invariance in a 2-tier HCN).** Consider a two tier network with BS activity factors $p_1$ and $p_2$. If $p_1 < p_2$, increasing the density of the first tier leads to a higher increase in the interference power due to the higher likelihood of having a closer tier-1 BS as the serving BS but a relatively smaller increase in the interference power. The coverage probability in this case is expected to increase. On the other hand, if $p_1 > p_2$, increasing the density of tier-1 BSs leads to higher increase in the interference power as compared to the intended power, leading to a decrease in the coverage probability. The two effects cancel each other when the activity factors of the two tiers are the same.

We now extend this result and derive exact condition under which the addition of $(K + 1)^{th}$ tier won’t affect (or will improve) the coverage of the existing $K$-tier network. We again assume same target SIR for all the tiers. The result is given in the following Corollary.

**Corollary 4 (Same $\beta$: Effect of adding $(K + 1)^{th}$ tier).** The overall coverage probability increases with the addition of the $(K + 1)^{th}$ tier if the load on the new tier satisfies:

$$P_{K+1} < \sum_{l=1}^{K} p_l \frac{\lambda_i P_i^{2/\alpha}}{\sum_{i=1}^{K} \lambda_i P_i^{2/\alpha}},$$

(28)

decreases if the inequality is reversed and remains the same if (28) holds with equality.

**Proof:** From Corollary 3 we note that the only term in the coverage probability expression that will change with the addition of a new tier is $A/\eta$. It can be expressed as:

$$A/\eta = \frac{\pi \Gamma(1 + \frac{2}{\alpha})}{C(\alpha) \beta^{2/\alpha}} \left[ \left( \frac{\sum_{i=1}^{K} p_i \lambda_i P_i^{2/\alpha}}{\sum_{i=1}^{K} \lambda_i P_i^{2/\alpha}} \right)^{-1} - 1 \right],$$

(29)

Defining effective load on a $K$-tier network as:

$$P_{\text{eff}}^{(K)} = \sum_{l=1}^{K} p_l \frac{\lambda_i P_i^{2/\alpha}}{\sum_{i=1}^{K} \lambda_i P_i^{2/\alpha}},$$

(30)
$A/\eta$ can be expressed as:

$$\frac{A}{\eta} = \frac{\pi \Gamma(1 + \frac{\alpha}{2})}{C(\alpha)\beta^{2/\alpha}} \left( 1 - p^{(K)}_{\text{eff}} \right),$$

(31)

which is the same as (23) for the single tier coverage result derived in Corollary 2. From this equivalence, it follows that
the coverage probability is a decreasing function of $p_{\text{eff}}$. Therefore, if the addition of the new tier leads to lower
effectiveness on the network, the coverage will increase. This
can be shown to be the case when (28) holds as follows:

$$p^{(K+1)}_{\text{eff}} = \sum_{i=1}^{K+1} \frac{\lambda_i P_i^{2/\alpha}}{\sum_{i=1}^{K+1} \lambda_i P_i^{2/\alpha}}$$

(32)

$$\leq \sum_{i=1}^{K} \frac{\lambda_i P_i^{2/\alpha}}{\sum_{i=1}^{K} \lambda_i P_i^{2/\alpha}} + \frac{\lambda_{K+1} P_{K+1}^{2/\alpha}}{\sum_{i=1}^{K} \lambda_i P_i^{2/\alpha}}$$

(33)

$$= p^{(K)}_{\text{eff}}.$$  

(34)

The other two results follow using the same argument. 

C. Bounds on the Coverage Probability

Evaluation of the exact expression of the coverage probability
requires an infinite summation. Although we have argued
that the summation can be tightly approximated by considering
only a first few terms, we haven’t yet provided a formal
method to determine the exact number of terms required such
that the approximation error is within predefined limit, say $\epsilon$.
Interestingly, this can be achieved as a by-product of the set
of bounds we derive in this section that can be made arbitrarily
tight. The idea is to use the following identity of $\exp(-x)$.

**Lemma 3.** For $x \geq 0$ and $m > 0$,

$$\sum_{i=0}^{2m-1} \frac{(-x)^i}{i!} \leq \exp(-x) \leq \sum_{i=0}^{2m} \frac{(-x)^i}{i!}.$$  

(35)

**Proof:** The proof follows from induction. Since $\exp(0) = 1$ and $\sum_{i=0}^{2m-1} \frac{(-x)^i}{i!} = 1$, it suffices to prove
that $\frac{d}{dx} \sum_{i=0}^{2m-1} \frac{(-x)^i}{i!} - \exp(-x) < 0$, which follows from the upper bound when $m = m - 1$.

Using this identity in the proof of Theorem 1 results in the
following bounds.

**Lemma 4 (Bounds on Coverage Probability).** For $m > 0$,
the coverage probability for the proposed load model can be bounded as

$$- \sum_{i=1}^{2m} g(i) \leq P_c - \frac{\pi}{C(\alpha)} \frac{\sum_{i=1}^{K} p_i \lambda_i P_i^{2/\alpha} \beta_i^{-2/\alpha}}{\sum_{i=1}^{K} \lambda_i P_i^{2/\alpha}} \leq - \sum_{i=1}^{2m-1} g(i)$$

(36)

Clearly, these bounds can be made arbitrarily tight by in-
creasing the value of $m$. Interestingly, these bounds are closely
related to the exact expression of coverage probability derived
in Theorem 1. In particular, the upper and lower bounds are
derived by truncating the infinite sum over $g(m)$ at odd and
even number of terms, respectively. Therefore, these bounds
provide a direct way to find the number of terms of $g(m)$
required to ensure an approximation error within a predefined
limit $\epsilon$, which is equal to $M_\epsilon$, where $M_\epsilon = \min \{ g(m) \} < \epsilon$.
We will use this observation in the study of the convergence
of infinite sum over $g(m)$ in the Numerical Results Section.

We conclude this section by noting that some terms of the
sequence $g(m)$ can be expressed in closed form, leading to
closed form bounds for the special case when $\alpha = 4$ and
$m = 2$. The bounds in this case depend only on the first two
terms of $g(m)$ that can be expressed as:

$$g(1) = \frac{-A}{\eta} \left\{ 2 \sqrt{\pi} \frac{\lambda_i p_i P_i^{1/2} \beta_i^{-1/2}}{1 + \sqrt{1 + \beta_i}} \right\}$$

(37)

$$g(2) = \left( \frac{A}{\eta} \right)^2 \left\{ 1 - \frac{2 \pi}{\eta} \sum_{i=1}^{K} \lambda_i p_i P_i^{1/2} (\beta_i^{-1/2} - \csc^{-1}(\sqrt{1 + \beta_i})) \right\}.$$  

(38)

IV. NUMERICAL RESULTS

Since most of the analytical results derived in this paper are
fairly self-explanatory and do not require separate numerical
treatment, we will provide only those results which help in
validating key modeling assumptions or help better visualize
certain important trends.

A. Convergence of Infinite Sum

We study the convergence of the infinite sum appearing in
the coverage probability expression in Figs. 2 and 3. Fig. 2
plots the truncated series $\sum_{i=1}^{m} g(i)$ as the function of $m$
for a single tier network and hence gives insights about the number
of terms required until the series converges. To understand
the trends, recall that the ratio $A/\eta$ decreases monotonically with
the activity factor $p$. Therefore, the number of terms required
for the series to converge are higher when the BS activity

![Fig. 2. Plot showing the convergence of the series $\sum_{i=1}^{m} g(i)$ for various BS activity factors in a single tier network with $\beta = 1$.](image-url)
factor is lower. Moreover, for the case when \( A/\eta > 1 \), i.e., \( p = .25 \), the series first increases until a certain point and then decreases and finally converges to its limiting value. This trend has been discussed in detail earlier in the paper when we proved the convergence of the infinite sum. To provide an idea of the number of terms required such that the approximation is within \( \epsilon \) of the exact value, we plot the number of terms \( M_\epsilon \) for various scenarios in Fig. 3. We again note that the number of terms required are reasonably small unless the transmission probability of some tier is extremely small.

**B. System Model Validation**

1) **Comparison with the fully loaded system:** After gaining insights into the behavior of the coverage probability expression, we now use it to highlight the importance of the proposed model by comparing the coverage results of a single tier network with those of a fully loaded system in Fig. 4. Although a huge difference in the coverage guarantees was expected for very low BS activity factors, it is indeed interesting that the coverage estimates assuming full load are quite pessimistic even for reasonably high load scenarios, such as \( p = .7 \) – .8.

2) **Comparison with an actual 4G deployment:** After highlighting the importance of the proposed load model, we now validate the PPP model used for the BS locations in this paper. While this model seems sensible for the small cells, especially the ones driven by unplanned user deployments, such as femtocells, it is dubious for the centrally planned tiers such as macrocells. Therefore, with a special focus on the macrocells, we consider three location models for a two tier HCN: i) macrocell locations modeled as a realization of a PPP, ii) macrocell locations modeled as a hexagonal grid, iii) macrocell locations drawn from an actual 4G deployment over 40 × 40 Km², \( \alpha = 3.8 \), \( p = [0.8, 0.6] \), \( \beta_1 = \beta_2 = \beta \).
importance of the proposed load model.

3) Comparison with a detailed simulation: The following two assumptions were made to facilitate analysis: i) the activity factors are the same for all the BSs of a particular tier, and ii) the activity of each BS is independent of the other BSs. We now validate these assumptions by comparing the analytical results with a detailed system simulation. For this comparison, we consider a simulation setup consisting of a two tier HCN, with the BSs of each tier modeled by independent PPPs. The user locations are also modeled by an independent PPP with density $\lambda_u$. As in the proposed model, each user is associated with the BS that provides the best received signal strength. From this, we calculate the actual load being served by each BS in terms of the number of users, which we denote by $N_{x_i}$ for a BS located at $x_i$. Assuming the number of orthogonal resource blocks, e.g., time-frequency resource blocks in LTE [20], to be $M$, the activity factor of a BS in each resource block can be expressed as $p_{x_i} = N_{x_i}/M$ as discussed in Section IV. To keep the setup simple, we consider the regime where the probability of having $N_{x_i} > M$ for any BS is small and whenever it happens, the activity factor for that BS is assumed to be 1. For this setup, the coverage probability results are presented in Fig. 6.

For a meaningful comparison of this simulation result with the analytical results, we first need to find analytical expressions of the activity factors $p_i$ as a function of the user density $\lambda_u$. For this, we leverage Corollary 2 of [17], where it is shown that the fraction of users served by $j^{th}$ tier is given by:

$$\bar{N}_j = \frac{\lambda_j (P_j/\beta_j) \frac{2}{P_j}}{\sum_{i=1}^{K} \lambda_i (P_i/\beta_i) \frac{2}{P_i}}.$$

(39)

Using this result, the average number of users served by a $j^{th}$ tier BS (average load) is $\lambda_u \bar{N}_j$. Therefore, assuming $M$ resource blocks, the activity factor in a randomly chosen resource block is:

$$p_j = \frac{1}{M} \frac{\lambda_u \bar{N}_j}{\lambda_j} = \frac{\lambda_u (P_j/\beta_j)^\frac{2}{P_j}}{M \sum_{i=1}^{K} \lambda_i (P_i/\beta_i)^\frac{2}{P_i}}.$$  

(40)

We use this analytical result for the load factors in the coverage probability results derived in the paper to plot the analytical results as a function of the user density in Fig. 6. Comparing this result with the numerical result obtained from a detailed simulation, we note that the two are reasonably close, especially relative to the previously known results for the fully-loaded system. This validates the two assumptions mentioned in the starting of this discussion.

C. Scale Invariance and Effect of Adding Small Cells

We now consider a two tier system and plot the coverage probability as a function of the density of second tier for various BS activity factors in Fig. 7. The target $\text{SIR}$ is fixed to be the same for both the tiers. We first note that the network is invariant to the changes in density when $p_1 = p_2$ as discussed in the last section. More importantly, we note that the coverage probability increases with $\lambda_2$ when the second tier BSs are less active than the first tier. This is an important result from the perspective of small cells, which are generally less active than macrocell BSs. Therefore, the coverage probability of the network should increase with the addition of small cells in this regime. This is a strong rebuttal to the viewpoint that unplanned infrastructure might bring down a cellular network due to increased interference. On the other hand, if a tier of BSs is added which is more active than the macrocells, the coverage would decrease, although this case seems pretty unlikely given the high load handled by the macrocells.

D. Open vs Closed Access

So far in this section we have only studied open access networks, where a mobile user can access any BS in the network. We now study the effect of closed access on the
coverage probability, with a focus on a particular scenario of interest where a certain fraction of BSs of a particular tier are in closed access while the others are in open access. This scenario is especially important in the current HCNs, where the access permissions of a particular small cell might be different for different set of users. It is easy to argue that this scenario can be visualized as a special case of the general model developed in this paper. To understand this, assume that a fraction $f$ of BSs of a certain tier are in open access – we assume that a BS is in open or closed access independent of the other BSs. Therefore, the density of BSs in open access $\lambda_i^{(o)}$ and closed access $\lambda_i^{(c)}$ can be evaluated from the following two equations:

$$f = \frac{\lambda_i^{(o)}}{\lambda_i^{(o)} + \lambda_i^{(c)}} \quad \text{(41)}$$

$$\lambda_i = \lambda_i^{(o)} + \lambda_i^{(c)} \quad \text{(42)}$$

Now this tier can be divided into two tiers, one with density $\lambda_i^{(o)}$, which is in open access, and other with density $\lambda_i^{(c)}$, which is in closed access – both form independent PPPs.

To study this scenario in detail, we consider a two tier HCN, where the first tier is in open access and fraction $1 - f$ of BSs of the second tier is in closed access. For this scenario, the coverage probability as a function of $f$ is presented in Fig. 8 for various load scenarios. The results confirm the intuition that the gap in open and closed access results reduces when the value of $f$ is increased. More interestingly, the gap is smaller when the second tier BSs are lightly loaded. This implies that the effect of interference due to closed access small cells on coverage probability is negligible if there are enough small cells in open access.

V. CONCLUSIONS

In this paper, we have incorporated a flexible notion of BS load in random spatial models for $K$-tier HCNs by introducing a new idea of conditionally thinning the interference field, conditional on the connection of a typical mobile to its serving BS. We have shown that this conditional thinning is a natural way of modeling different levels of load on different types of BSs arising mainly from the differences in their coverage footprints. We observe that the fully loaded models are extremely pessimistic in terms of coverage, and the analysis shows that adding lightly loaded access points (e.g. pico or femtocells) to the macrocell network always increases coverage probability.

This work has numerous extensions. Firstly, it is important to exactly model the temporal and spatial correlation in the BS activity factors and develop tools to incorporate it in the framework developed in this paper. Secondly, it is important to understand the coverage and rate trends for this conditional-thinning using more realistic spatial models that include inter-point interactions, such as Gibbs models [27]. Another related idea is to define BS load from queuing perspective and then study spatial and temporal dynamics simultaneously. Since this work focused only on the downlink, an interesting extension is to develop a similar framework for uplink possibly using some of the ideas recently developed in [33].

APPENDIX A

PROOF OF LEMMA 2

Being consistent with the definition of $T$, we note that:

$$T^m = 1 \left( \max_{i \in K} \frac{M(\Psi_i)}{\delta_i} < I \right) I^{-2m/\alpha} \quad \text{(43)}$$

To proceed with the proof, we represent $I^{-2m/\alpha}$ in terms of $\Gamma(x)$ as:

$$I^{-2m/\alpha} = \frac{1}{\Gamma(2m/\alpha)} \int_0^\infty e^{-sI}s^{-1 + \frac{2m}{\alpha}} ds, \quad m \geq 1, \quad \text{(44)}$$

where $\Gamma(x)$ is the standard gamma function. Using this representation of $I^{-2m/\alpha}$ we can express $\mathbb{E}[T^m]$ as:

$$\mathbb{E}\left[ 1 \left( \max_{i \in K} \frac{M(\Psi_i)}{\delta_i} < I \right) \frac{1}{\Gamma(2m/\alpha)} \int_0^\infty e^{-sI}s^{-1 + \frac{2m}{\alpha}} ds \right] \quad \text{(45)}$$

Using Fubini’s theorem, we can exchange the expectation and the inner integral to obtain

$$\frac{1}{\Gamma(2m/\alpha)} \int_0^\infty s^{-1 + \frac{2m}{\alpha}} \mathbb{E} \left[ e^{-sI} \mathbf{1} \left( \max_{i \in K} \frac{M(\Psi_i)}{\delta_i} < I \right) \right] ds. \quad \text{(46)}$$

Under the assumption $\beta_i > 1, \forall i$, we know that only one BS in the whole network can establish a downlink connection with a typical mobile. Hence,

$$\mathbb{E} \left[ e^{-sI} \mathbf{1} \left( \max_{i \in K} \frac{M(\Psi_i)}{\delta_i} < I \right) \right] = \sum_{i=1}^K \sum_{x \in \Psi_i} \mathbf{1} (\text{SIR}(x) > \beta_i), \quad \text{(47)}$$

where SIR(x) is the received SIR when a typical mobile is camped to the BS located at $x \in \Psi_i$. Using this expression, the expectation term of (46) can be written as:

$$\mathbb{E} \left[ e^{-sI} \mathbf{1} \left( \max_{i \in K} \frac{M(\Psi_i)}{\delta_i} < I \right) \right]$$

Fig. 8. Coverage probability in a two tier network as a function of the fraction of the second tier BSs in open access ($\beta = [1, 1], P = [1, 0.1], \lambda_2^{(c)} = 10\lambda_1^{(c)}, p_1 = 1$ and $\alpha = 3.8$). The density of second tier BSs in open access $\lambda_2^{(o)} = \frac{f}{1-f} \lambda_2^{(c)}$, where $f$ is the fraction of BSs in open access.
\begin{align*}
&= \mathbb{E}\left[ e^{-sl} \right] - \sum_{i=1}^{K} \mathbb{E}\left[ e^{-sl} \sum_{x \in \Psi} \mathbf{1}(\text{SIR}(x) > \beta_i) \right]. \\
\text{(48)}
\end{align*}

From Lemma [1], we know the Laplace transform of total interference and hence the first term in the above expression can be directly written as:
\begin{align*}
\mathbb{E}\left[ e^{-sl} \right] &= \exp(\lambda^{2/\alpha}) C(\alpha) \sum_{i=1}^{K} p_i \lambda_i P_i^{2/\alpha}.
\text{(49)}
\end{align*}

To evaluate the expectation in the second term of (48), we first denote the effective interference as \( I' = I - P_i h_x \|x\|^{-\alpha} \) and note that the Laplace transforms of \( I \) and \( I' \) are the same. The expectation can now be simplified as:
\begin{align*}
\mathbb{E}\left[ e^{-sl} \sum_{x \in \Psi} \mathbf{1}(\text{SIR}(x) > \beta_i) \right] &= \mathbb{E}\sum_{x \in \Psi} \exp(-sl' + P_i h_x \|x\|^{-\alpha}) \mathbf{1}(\frac{P_i h_x \|x\|^{-\alpha}}{I'} > \beta_i) \\
\text{(a)} &= \mathbb{E}\left[ \sum_{x \in \Psi} e^{-sl'} \mathbf{1}(h_x > \beta_i I' P_i^{-1} \|x\|^\alpha) \right] \\
\text{(b)} &= \mathbb{E}\left[ \sum_{x \in \Psi} E_P\left[ \exp(-I'(s(1 + \beta_i) + \beta_i P_i^{-1} \|x\|^\alpha)) \right] \right], \\
\text{(50)}
\end{align*}

where (a) follows from the fact that fading is independent of all the other random variables and (b) follows from the fact that \( h_x \sim \exp(1) \). Now, using the Laplace transform of \( I' \) and recalling \( \eta = C(\alpha) \sum_{i=1}^{K} \lambda_i \lambda_i P_i^{2/\alpha} \), it can be further simplified to:
\begin{align*}
\mathbb{E}\left[ \sum_{x \in \Psi} \exp(-\eta(s(1 + \beta_i) + \beta_i P_i^{-1} \|x\|^\alpha) \frac{\lambda^{2/\alpha}}{1 + s P_i \|x\|^{-\alpha}}) \right], \\
\text{(53)}
\end{align*}

and using Campbell Mecke theorem [13] to:
\begin{align*}
\lambda_i P_i \int_{\mathbb{R}^2} (s P_i)^{-1/\alpha} e^{-s P_i \|x\|^{-\alpha}} \eta^{m-1} \alpha(m-1)! \frac{\lambda^{2/\alpha}}{1 + s P_i \|x\|^{-\alpha}} \frac{d\eta(x)}{s}.
\text{(54)}
\end{align*}

With this we have now simplified both the terms of (48) given respectively by (49) and (54). We now substitute the first term in (46) and evaluate the integral with respect to \( s \) as:
\begin{align*}
\int_{0}^{\infty} s^{-1+2m/\alpha} \mathop{\exp\left( -\eta s^{2/\alpha} \right)} dx = \frac{\eta^{-m} \alpha(m-1)!}{2}, \\
\text{(55)}
\end{align*}

where the solution follows from the substitution \( s^{2/\alpha} \rightarrow y \) followed by integration by parts. Now substituting the second term (given by (53) in (46), we get the following integral:
\begin{align*}
\lambda_i P_i \int_{0}^{\infty} \int_{\mathbb{R}^2} s^{-1+2m/\alpha} e^{-s(1 + \beta_i) + \beta_i P_i^{-1} \|x\|^\alpha} \frac{\lambda^{2/\alpha}}{1 + s P_i \|x\|^{-\alpha}} d\eta(x) d\eta(y).
\text{(56)}
\end{align*}

Now use the substitution \( (s P_i)^{-1/\alpha} x \rightarrow x \), which leads to
\begin{align*}
\lambda_i P_i \int_{\mathbb{R}^2} s^{-1+2m/\alpha} e^{-\eta \lambda^{2/\alpha}(1 + \beta_i) + \beta_i \|x\|^\alpha} \frac{\lambda^{2/\alpha}}{1 + \|x\|^{-\alpha}} (s P_i)^{-\alpha} d\eta(x).
\text{(57)}
\end{align*}

Now exchange the integrals to obtain
\begin{align*}
\lambda_i P_i \int_{\mathbb{R}^2} \int_{0}^{\infty} s^{-1+2(m+1)/\alpha} e^{-\eta \lambda^{2/\alpha}(1 + \beta_i) + \beta_i \|x\|^\alpha} \frac{\lambda^{2/\alpha}}{1 + \|x\|^{-\alpha}} ds dx.
\text{(58)}
\end{align*}

Now the inner integral (with respect to \( s \)) can be evaluated directly using the definition of \( \Gamma(x) \) function or using the substitution \( s^{2/\alpha} \rightarrow s \) to obtain the below integral.
\begin{align*}
\lambda_i P_i \int_{\mathbb{R}^2} \frac{\lambda^{2/\alpha} \alpha m!}{2\eta^{m-1}} s^{\alpha} \frac{d\eta(x)}{(1 + \|x\|^{-\alpha})(1 + \beta_i + \beta_i \|x\|^\alpha)^{(m+1)/\alpha}(m+1)^{1/\alpha}}.
\text{(59)}
\end{align*}

Now the above integral can be expressed as:
\begin{align*}
\frac{1}{(1 + \beta_i)^{(m+1)/\alpha}} \int_{\mathbb{R}^2} \frac{d\eta(x)}{(1 + \|x\|^{-\alpha})(1 + \beta_i + \beta_i \|x\|^\alpha)^{(m+1)/\alpha}(m+1)^{1/\alpha}}.
\text{(60)}
\end{align*}

Now using the substitution \( 1 + \beta_i \|x\|^\alpha \rightarrow t^{-1} \), the above expression can be simplified to
\begin{align*}
\frac{2\pi \beta_i^{-2/\alpha}}{\Gamma(2m/\alpha) \Gamma(1 + 2/\alpha)} \frac{\Gamma(1 + (m + 1)/\alpha)}{\Gamma(1 + (m + 1)/\alpha)} \frac{\Gamma(2m/\alpha) \Gamma(1 + 2/\alpha)}{\Gamma(1 + (m + 1)/\alpha)} 2F_1(1, 2m/\alpha, 1 + (m + 1)/2/\alpha, (1 + \beta_i)^{-1}),
\text{(61)}
\end{align*}

where \( 2F_1 \) is the generalized hypergeometric function. Combining all the above we obtain the result.

\begin{flushright}
\text{\small ■}
\end{flushright}

\begin{thebibliography}{11}
\bibitem{1} H. S. Dhillon, R. K. Ganti, and J. G. Andrews, “Load-aware heterogeneous cellular networks: modeling and SIR distribution,” in \textit{Proc. 2012 IEEE GLOBECOM}.
\bibitem{2} Cisco, “Cisco visual networking index: global mobile data traffic forecast update, 2011-2016,” white paper, Feb. 2012.
\bibitem{3} Qualcomm, “LTE advanced: heterogeneous networks,” white paper, Jan. 2011.
\bibitem{4} J. G. Andrews, H. Claussen, M. Dopfer, S. Rangan, and M. C. Reed, “Femtocells: past, present, and future,” \textit{IEEE J. Sel. Areas Commun.}, vol. 18, no. 11, pp. 2443–2455, Nov. 2000.
\bibitem{5} S. Kishore, L. Greenstein, H. Poor, and S. Schwartz, “Uplink user capacity in a CDMA macrocell with a hotspot microcell: exact and approximate analyses,” \textit{IEEE Trans. Wireless Commun.}, vol. 2, no. 2, pp. 364–374, Mar. 2003.
\bibitem{6} W. Roh and A. Paulraj, “Performance of the distributed antenna systems in a multi-cell environment,” in \textit{Proc. 2003 IEEE Veh. Technology Conf. – Spring}, pp. 587–591.
\bibitem{7} T. S. Rappaport, \textit{Wireless Communications: Principles and Practice}, 2nd edition. Prentice-Hall, 2002.
\bibitem{8} A. Wyner, “Shannon-theoretic approach to a Gaussian cellular multiplex-access channel,” \textit{IEEE Trans. Inf. Theory}, vol. 40, no. 6, pp. 1713–1727, Nov. 1994.
\bibitem{9} T. X. Brown, “Cellular performance bounds via shotgun cellular systems,” \textit{IEEE J. Sel. Areas Commun.}, vol. 18, no. 11, pp. 2443–2455, Nov. 2000.
\bibitem{10} F. Baccelli and S. Zuyev, “Stochastic geometry models of mobile communication networks,” in \textit{Frontiers in Queueing}. CRC Press, 1997, pp. 227–243.
\bibitem{11} J. G. Andrews, F. Baccelli, and R. K. Ganti, “A tractable approach to coverage and rate in cellular networks,” \textit{IEEE Trans. Commun.}, vol. 59, no. 11, pp. 3122–3134, Nov. 2011.
\bibitem{12} P. J. Fleming, A. L. Stolyar, and B. Simon, “Closed-form expressions for other-cell interference in cellular CDMA,” Technical Report 1116, Univ. of Colorado at Boulder, Dec. 1997.
\bibitem{13} D. Stoyan, W. S. Kendall, and J. Mecke, \textit{Stochastic Geometry and Its Applications}, 2nd edition. John Wiley and Sons, 1995.
\bibitem{14} J. F. C. Kingman, \textit{Poisson Processes}. Oxford University Press, 1993.
\bibitem{15} F. Baccelli and B. Blaszczyszyn, \textit{Stochastic Geometry and Wireless Networks}. NOW: Foundations and Trends in Networking, 2009.
\end{thebibliography}
H. S. Dhillon, R. K. Ganti, and J. G. Andrews, “A tractable framework for coverage and outage in heterogeneous cellular networks,” in Proc. 2011 Information Theory and its Applications.

H. S. Dhillon, R. K. Ganti, F. Baccelli, and J. G. Andrews, “Modeling and analysis of K-tier downlink heterogeneous cellular networks,” IEEE J. Sel. Areas Commun., vol. 30, no. 3, pp. 550–560, Apr. 2012.

K. Mulkerjee, “Distribution of downlink SIR in heterogeneous cellular networks,” IEEE J. Sel. Areas Commun., vol. 30, no. 3, pp. 575–585, Apr. 2012.

H.-S. Jo, Y. J. Sang, P. Xia, and J. G. Andrews, “Heterogeneous cellular networks with flexible cell association: a comprehensive downlink SINR analysis,” IEEE Trans. Wireless Commun., vol. 11, no. 10, pp. 3484–3495, Oct. 2012.

H. S. Dhillon, R. K. Ganti, F. Baccelli, and J. G. Andrews, “Coverage and ergodic rate in K-tier downlink heterogeneous cellular networks,” in Proc. 2011 Allerton Conf. on Comm., Control, and Computing.

A. Dammjanovic, J. Montojo, Y. Wei, T. Ji, T. Luo, M. Vajapeyam, T. Yoo, O. Song, and D. Malladi, “A survey on 3GPP heterogeneous networks,” IEEE Wireless Commun. Mag., vol. 18, no. 3, pp. 10–21, Jun. 2011.

A. Ghosh, J. G. Andrews, N. Mangalvedhe, R. Ratasuk, B. Mondal, M. Cudak, E. Visotsky, T. A. Thomas, P. Xia, H. S. Jo, H. S. Dhillon, and T. D. Novlan, “Heterogeneous cellular networks: from theory to practice,” IEEE Commun. Mag., Jun. 2012.

B. Rengarajan and G. de Veciana, “Architecture and abstractions for environment and traffic-aware system-level coordination of wireless networks,” IEEE/ACM Trans. Networking, vol. 19, no. 3, pp. 721–734, Jun. 2011.

S. Borst, N. Hegde, and A. Proutièrè, “Interacting queues with server selection and coordinated scheduling—application to cellular data networks,” Ann. Oper. Res., vol. 170, pp. 59–78, 2009.

S. Borst, “User-level performance of channel-aware scheduling algorithms in wireless data networks,” in Proc. 2003 IEEE INFOCOM, pp. 321–331.

T. Bonald and A. Proutièrè, “Wireless downlink data channels: user performance and cell dimensioning,” in Proc. 2003 ACM MobiCom.

D. B. Taylor, H. S. Dhillon, T. D. Novlan, and J. G. Andrews, “Pairwise interaction processes for modeling cellular network topology,” in Proc. 2012 IEEE Globecom.

G. Boudreau, J. Panicker, N. Guo, R. Chang, N. Wang, and S. Vzic, “Interference coordination and cancellation for 4G networks,” IEEE Commun. Mag., vol. 47, no. 4, pp. 74–81, Apr. 2009.

F. Baccelli, P. Muhlethaler, and B. Blaszczyszyn, “Stochastic analysis of spatial and opportunistic Aloha,” IEEE J. Sel. Areas Commun., vol. 27, no. 7, pp. 1105–1119, Sep. 2009.

A. Ghosh, J. Zhang, J. G. Andrews, and R. Muhamed, Fundamentals of LTE, Prentice-Hall, 2010.

R. Ganti and M. Haenggi, “Spatial and temporal correlation of the interference in ALOHA ad hoc networks,” IEEE Commun. Lett., vol. 13, no. 9, pp. 631–633, Sep. 2009.

U. Schlicker, C. Bettstetter, and G. Brandner, “Temporal correlation of interference in wireless networks with Rayleigh block fading,” IEEE Trans. Mobile Computing, vol. 11, no. 12, pp. 2109–2120, Dec. 2012.

T. D. Novlan, H. S. Dhillon, and J. G. Andrews, “Analytical modeling of uplink cellular networks,” submitted to IEEE Trans. Wireless Commun., Mar. 2102. Available: arxiv.org/abs/1203.1304.

Harpreet S. Dhillon (S’11) received the B.Tech. degree in Electronics and Communication Engineering from IIT Guwahati, India, in 2008 and the M.S. in Electrical Engineering from Virginia Tech in 2010. He is currently a Ph.D. student at The University of Texas at Austin, where his research has focused on the modeling and analysis of heterogeneous cellular networks using tools from stochastic geometry, point process theory and spatial statistics. His other research interests include interference channels, multi-user MIMO systems and cognitive radio networks. He is the recipient of the Microelectronics and Computer Development (MCD) fellowship from UT Austin and was also awarded the Agilent Engineering and Technology Award 2008. He has held summer internships at Alcatel-Lucent Bell Labs in Crawford Hill, NJ, Samsung Dallas Technology Lab in Richardson, TX, Qualcomm Inc. in San Diego, CA, and Cercom, Politecnico di Torino in Italy.

Radha Krishna Ganti (S’01, M’10) is an Assistant Professor at the Indian Institute of Technology Madras, Chennai, India. He was a Postdoctoral researcher in the Wireless Networking and Communications Group at UT Austin from 2009-11. He received his B. Tech. and M. Tech. in EE from the Indian Institute of Technology, Madras, and a Masters in Applied Mathematics and a Ph.D. in EE form the University of Notre Dame in 2009. His doctoral work focused on the spatial analysis of interference networks using tools from stochastic geometry. He is a co-author of the monograph Interference in Large Wireless Networks (NOW Publishers, 2008).

Jeffrey G. Andrews (S’98, M’02, SM’06, F’13) received the B.S. in Engineering with High Distinction from Harvey Mudd College in 1995, and the M.S. and Ph.D. in Electrical Engineering from Stanford University in 1999 and 2002, respectively. He is a Professor in the Department of Electrical and Computer Engineering at the University of Texas at Austin, where he was the Director of the Wireless Networking and Communications Group (WNCG) from 2008-12. He developed Code Division Multiple Access systems at Qualcomm from 1995-97, and has consulted for entities including Verizon, the WiMAX Forum, Intel, Microsoft, Apple, Samsung, Clearwire, Sprint, and NASA. He is a member of the Technical Advisory Boards of Accelera and Fastback Networks.

Dr. Andrews is co-author of two books, Fundamentals of WiMAX (Prentice-Hall, 2007) and Fundamentals of LTE (Prentice-Hall, 2010), and holds the Earl and Margaret Brasfield Endowed Fellowship in Engineering at UT Austin, where he received the ECE department’s first annual High Gain award for excellence in research. He is a Senior Member of the IEEE, a Distinguished Lecturer for the IEEE Vehicular Technology Society, served as an associate editor for the IEEE TRANSACTIONS ON WIRELESS COMMUNICATIONS from 2004-08, was the Chair of the 2010 IEEE Communication Theory Workshop, and is the Technical Program co-Chair of ICC 2012 (Comm. Theory Symposium) and Globecom 2014. He is an elected member of the Board of Governors of the IEEE Information Theory Society and an IEEE Fellow.

Dr. Andrews received the National Science Foundation CAREER award in 2007 and has been co-author of five best paper award recipients, two at Globecom (2006 and 2009), Asilomar (2008), the 2010 IEEE Communications Society Best Tutorial Paper award, and the 2011 Communications Society Heinrich Hertz Prize. His research interests are in communication theory, information theory, and stochastic geometry applied to wireless cellular and ad hoc networks.