A retrospective study on Lindley distribution

Abstract

Generalizations make a distribution more flexible especially for studying the tail properties. There exist many generalized family of continuous univariate distributions. In this article, a survey on Lindley distribution, its extensions and classes is conducted. Several available generalizations of the distribution are reviewed and recent trends in the construction of generalized classes are discussed.

Keywords: generalizations, lindley distribution, power distribution, truncated distributions, wrapped distributions

Introduction

The analysis and modeling of lifetime data are very important in applied sciences such as engineering, public health, actuarial science, biomedical studies, demography, industrial reliability and other applied sciences. Therefore, it seems crucial to find statistical distributions for real–world phenomena. There are a number of lifetime distributions in statistical literature including exponential, Weibull, gamma and lognormal distributions. The Weibull distribution is one of the most popular and widely used models in life testing and reliability theory.1 Introduced a one–parameter distribution known as Lindley distribution. This model is used as an alternative model for reliability theory.

Generalizations of existing distribution is another interest in statistical research. The extended distributions have attracted in statistical literature to develop new models. The transformations of distributions have been proved useful in exploring skewness and tail properties, and also for improving the goodness–of–fit of the extended family. The Lindley distribution has been generalized by many authors in recent years. This study concentrates on conducting an extensive enquiry on those different fields of existing knowledge.

This paper is organized as follows. Section 2 deals with the Lindley Distribution. Section 3 discusses existing generalizations of Lindley distribution. Finally, concluding remarks are given in Section 4.

Lindley distribution

The probability density function (pdf) of one parameter Lindley distribution is given by

\[ f(x;\theta) = \frac{\theta^2}{1+\theta} (1+x)e^{-\theta x}; \quad x > 0, \theta > 0 \] (2.1)

The cumulative density function (cdf) of one parameter Lindley distribution, corresponding to the pdf given in equation (2.1) is

\[ F(x;\theta) = 1 - \left[ \frac{1}{1+\theta} e^{-\theta x} \right] e^{-\theta x}; \quad x > 0, \theta > 0 \] (2.2)

This distribution is derived as a mixture of exponential (\( \theta \)) and Gamma (2, \( \theta \)) distribution.

Hence the pdf takes the alternate form,

\[ f(x;\theta) = p f^1(x) + (1 - p) f^2(x) \]

where \( p = \frac{\theta^2}{1+\theta}, \quad f^1(x) = \theta e^{-\theta x}, \quad f^2(x) = \frac{\theta^2}{1+\theta} e^{-\theta x} x^{2-1} \)

and proved that in modeling and analysis of lifetime data, Lindley distribution provides a better model in many ways than the very usual exponential distribution. Along with various other properties they derived the central moments of Lindley distribution as,

\[ \mu_2 = \frac{\theta^2 + 4\theta + 2}{\theta^2(\theta + 1)^2} \]

and

\[ \mu_3 = \frac{2(\theta^3 + 6\theta^2 + 6\theta + 2)}{\theta^2(\theta + 1)^3} \]

and

\[ \mu_4 = \frac{3(3\theta^4 + 24\theta^3 + 44\theta^2 + 32\theta + 8)}{\theta^2(\theta + 1)^4} \]

Mazucheli discussed the applications of Lindley distribution on lifetime data regarding competing risks. Alternately, Shanker et al. made a comparison study of the goodness of fit of exponential and Lindley distributions on modeling of lifetime data. They provide different graphs for pdfs and cdfs for the same values of parameter for a visual comparison on the nature of the two distributions. They provide different associated functions as presented in Table 1.

The study associated fifteen different data sets and the fact revealed is that in some cases exponential distribution provides better fit than the Lindley distribution whereas in other cases Lindley distribution provides better fit than the exponential distribution. Thus does not make a final conclusion on the superiority of the two distributions. They arrive at a statement that the suitability depends on the nature of data. Exponential is simple, still Lindley is more flexible.

Quasi Lindley distribution

Shanker introduced Quasi Lindley distribution (QLD). QLD with parameters \( \alpha \) and \( \theta \) is defined by its pdf

\[ f(x,\alpha,\theta) = \frac{\theta(\alpha + x\theta)}{\alpha + 1} e^{-\theta x}, \quad x > 0, \theta > 0, \alpha > -1 \] (2.3)

It can be easily seen that at \( \alpha = \theta \), the QLD equation (2.3) reduces to the Lindley distribution given by equation (2.1) and at \( \alpha = 0 \), it reduces to the gamma distribution with parameters (2, \( \theta \)). The pdf equation (2.3) can be shown as a mixture of exponential (\( \theta \)) and gamma (2, \( \theta \)) distributions as follows:

\[ f(x,\alpha,\theta) = p f^1(x) + (1 - p) f^2(x) \]
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where \( p = \frac{a}{a+1} \), \( f_1(x) = \theta e^{-\theta x} \) and \( f_2(x) = \theta^2 x e^{-\theta x} \). The cdf of the QLD is obtained as

\[
F(x, \alpha, \theta) = 1 - \frac{1 + a + \theta x}{a+1} e^{-\theta x}, \ x > 0, \theta > 0, a > -1
\]

Table 1 Different measures of exponential and Lindley distributions

| Measure                        | Exponential distribution | Lindley distribution |
|-------------------------------|--------------------------|----------------------|
| coefficient of variation      | 1                        | \( \frac{\theta^2 + 4\theta + 2}{\theta + 2} \) |
| coefficient of Skewness (\( \beta_1 \)) | 2                        | \( 2(\theta^3 + 6\theta^2 + 6\theta + 2) \) \( (\theta^2 + 4\theta + 2)^{\frac{3}{2}} \) |
| coefficient of Kurtosis (\( \beta_2 \)) | 9                        | \( 3(3\theta^4 + 24\theta^3 + 44\theta^2 + 32\theta + 8) \) \( (\theta^2 + 4\theta + 2)^2 \) |
| Index of dispersion, \( \gamma \) | \( \frac{\mu_2}{\mu_1} \) | \( \frac{\theta^2 + 4\theta + 2}{\theta(\theta^2 + 2\theta + 2)} \) |
| hazard rate function           | 0                        | \( \frac{\theta^2 (1 + x)}{1 + \theta + \theta x} \) |
| mean residual life function    | 1                        | \( \frac{\theta + 2 + \theta x}{\theta(1 + \theta + \theta x)} \) |

The QLD has been fitted to a number of data sets to which earlier the Lindley distribution has been fitted by others and it was found that almost all these data-sets, the QLD provides closer fits than those by the Lindley distribution.

Generalizations of Lindley distribution

Shanker et al.\(^4\) have comparative study on modeling of lifetime data using one parameter\(^5\) distribution and exponential distribution. Lindley distribution is not suitable for modeling data sets where there is large right tail or the tail approaches to zero at a faster rate. Such data sets are quite common in insurance problems and count data example in biology. This critical observation has motivated the authors to search for extensions of Lindley distribution. There exist many versions of generalized Lindley distribution and extensions. In this section, important among them are considered one by one.

Two parameter Lindley distribution

Shanker\(^6\) suggested a two parameter Lindley distribution (TPLD), of which the Lindley distribution is a particular case. There exist two forms type 1 and type 2. The pdf of type 1 is defined as

\[
f(x; \alpha, \theta) = \frac{\theta^2}{\alpha^2 + 1} (a + x) e^{-\theta x}; x > 0, \theta > 0, a \theta > -1
\]

When \( a = 0 \), it gives the Lindley distribution and \( a = 0 \), it gives the gamma (2, \( \theta \)) distribution. The pdf is a mixture of exponential \( \theta \) and gamma (2, \( \theta \)) distribution where the mixing constant is \( \frac{a \theta}{\alpha^2 + 1} \).

The cdf of type 1 TPLD is

\[
F(x; \alpha, \theta) = \frac{1 - \frac{1 + a + \theta x}{\alpha^2 + 1} e^{-\theta x}}{1 + \frac{a + \theta x}{\alpha^2 + 1} e^{-\theta x}}; x > 0, \theta > 0, a \theta > -1
\]

Then the corresponding failure rate function \( h(x) \) is

\[
h(x) = \frac{\theta^2 (1 + a \theta x)}{\theta + a + \theta a x}; x > 0, \theta > 0
\]

The pdf of type 2 is defined as

\[
f(x; \beta, \beta) = \frac{\theta^2}{\beta^2 + 1} (1 + \theta x) e^{-\theta x}; x > 0, \theta, \beta > 0
\]

The cdf of type 2 TPLD is,

\[
F(x; \beta, \beta) = 1 - \frac{\theta + \beta + \theta \beta x}{\beta} e^{-\theta x}; x > 0, \theta > 0, \beta > -\theta
\]

Then the corresponding failure rate function \( h(x) \) is

\[
h(x) = \frac{\theta^2 (1 + \theta x)}{\theta + \beta + \theta \beta x}; x > 0, \theta > 0, \beta > -\theta
\]

The mean of TPLD is always greater than the mode, the distribution is positively skewed. The TPLD provides better fits than those by the one parameter Lindley distribution.\(^7\) Obtained a two parameter weighted Lindley distribution and studied its applications to survival data.\(^8\) Studied two parameter Lindley distribution for modeling waiting and survival data.\(^9\) Proposed an extended Lindley distribution which offers a more flexible model for lifetime data.

Akash distribution

The one parameter lifetime distribution with pdf

\[
f(x) = \frac{\theta^2}{\theta^2 + 2} (1 + x^2) e^{-\theta x}; x > 0, \theta > 0
\]

Suggested by Shanker\(^10\) is termed as Akash distribution. The pdf is a mixture of exponential (\( \theta \)) and gamma (3, \( \theta \)) distribution where the mixing constant \( \frac{\theta^2}{\theta^2 + 2} \). The cdf of Akash distribution is

\[
F(x) = 1 - \frac{1 + \theta x}{\theta + \theta x} e^{-\theta x}, \ x > 0, \theta > 0
\]

\[
h(x) = \frac{\theta^2 (1 + \theta x)}{\theta + \theta x}; x > 0, \theta > 0
\]
A new generalized Lindley distribution

Ibrahim et al.\textsuperscript{14} proposed a new generalized Lindley distribution (NGLD) is obtained from a mixture of the gamma ($\beta, \theta$) and gamma ($\beta, \theta$) where the mixing constant is $\frac{1}{1+\theta}$. The pdf of NGLD is

$$f(x; \theta, \alpha, \beta) = \frac{1}{1+\theta} \left[ \frac{\theta^{\alpha+1}x^{\alpha-1}}{\Gamma(\alpha)\Gamma(\beta)} + \frac{\theta^\beta x^{\beta-1}}{\Gamma(\beta)} \right] e^{-\theta x}; \alpha, \theta > 0, x > 0$$

The corresponding cdf is given by

$$F(x; \theta, \alpha, \beta) = \frac{1}{1+\theta} \left[ \frac{\theta^\beta (\alpha+1)}{\Gamma(\alpha)\Gamma(\beta)} + \frac{\theta^\beta x^\beta}{\Gamma(\beta)} \right] e^{-\theta x}; \alpha, \theta > 0, x > 0$$

The hazard rate function $h(x)$ is

$$h(x) = \frac{1}{1+\theta} \left[ \frac{\theta^\beta (\alpha+1)}{\Gamma(\alpha)\Gamma(\beta)} + \frac{\theta^\beta x^\beta}{\Gamma(\beta)} \right] e^{-\theta x}; \alpha, \theta > 0, x > 0$$

The hazard rate function of GLD, is increasing for $\gamma > 0$, bathtub shaped for $\gamma > 0$ and $\gamma > 0$, decreasing for $\alpha \leq 1$ and $\alpha = 1$. At $\alpha = 1$ and $\beta = 2$, the NGLD becomes Lindley distribution. For $\alpha = \beta = \lambda$, NGLD becomes gamma distribution with parameter ($\theta, \lambda$) and at $\alpha = \beta = 1$ it becomes exponential distribution with parameter $\theta$. Gupta\textsuperscript{13} studied generalized exponential distribution. The Lindley distribution has been generalized by different researchers including.\textsuperscript{16–21}

Power Lindley distribution

Ghitany et al.\textsuperscript{22} introduced Power Lindley distribution (PLD) with pdf

$$f(x; \theta, \alpha) = \frac{\theta^\alpha}{\alpha^{\alpha+1}} (1+x^\alpha)^{\alpha-1} e^{-\theta x}; \theta, \alpha, x > 0$$

It is a new extension of Lindley distribution by considering the power transformation of the random variable $X = \frac{1}{y}$, where $Y$ follows Lindley distribution. At $\alpha = 1$ it reduces to Lindley distribution. It is a mixture of Weibull distribution ($2\alpha, \theta$) and generalized gamma distribution ($2\alpha, \theta$) with mixing proportion $\frac{\theta}{\theta+1}$. The corresponding cdf is given by

$$F(x; \theta, \alpha) = 1 - \frac{\theta+1+\theta x^\alpha}{\theta+1} e^{-\theta x}; \theta, \alpha, x > 0$$

The associated hazard rate function is

$$h(x) = \frac{\theta^\alpha x^{\alpha-1}(1+x^\alpha)}{\theta+1+\theta x^\alpha}; \theta, \alpha, x > 0$$

Extended power Lindley distribution

Alkarni\textsuperscript{23} suggested an Extended Power Lindley Distribution with parameters $\theta, \alpha$ and $\beta$ is defined by its pdf

$$f(x; \theta, \beta, \alpha) = \frac{\theta^\alpha}{\theta+\beta}(1+\beta x^\alpha)^{\alpha-1} e^{-\theta x}; \theta, \beta, \alpha, x > 0$$

It is a type 2 TPLD by considering the power transformation of the random variable $X = Y^{\frac{1}{\beta}}$, where $Y$ follows

$$F(x; \theta, \beta, \alpha) = 1 - \frac{\theta+1+\theta x^\alpha}{\theta+1+\theta x^\alpha}; \theta, \beta, \alpha, x > 0$$

The associated hazard rate function is

$$h(x) = \frac{\theta^\alpha x^{\alpha-1}(1+\beta x^\alpha)}{\theta+1+\beta x^\alpha}; \theta, \beta, \alpha, x > 0$$

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At α = β = 1. It gives the Lindley distribution, at α = 1 it gives TPLD, at β = 1 it gives power Lindley distribution and at γ = 0, it gives Weibull distribution.

The associated hazard rate function is

\[ h(x) = \frac{\alpha^2 (1+x)^{\beta-1} e^{-\theta x}}{\theta + x(1+\theta) e^{-\theta x}}, \quad x > 0, \alpha, \beta, \theta, \gamma > 0 \]

The coefficient of determination of Exponentiated Power Lindley Distribution is 0.975, which is higher than the coefficient of determination of PLD, GLD, Lindley Distribution, Exponentiated gamma distribution, and modified Weibull distributions. The Exponentiated Power Lindley Distribution is a good model for life time data.

The power Lindley distribution with its inference was proposed by Ghitany et al.\(^{22}\) and generalized by Liyanage\(^{18}\). Ghitany et al.\(^{22}\) discussed the estimation of the reliability of a stress–strength system by using transmuted Rayleigh distribution. Sharma et al.\(^{28}\) introduced the ILD as a one parameter model for a stress–strength reliability model. Alkarni\(^{28}\) proposed three parameter ILD with application to maximum flood level data. Also\(^{29}\) studied the extension of ILD.

### Truncated Lindley distribution

Ahmed et al.\(^{28}\) introduced truncated versions of Lindley Distribution. The truncated distribution is used where a random variable is restricted to be observed on some range. The truncated versions of the Lindley distribution, named as the upper truncated Lindley (UTL), lower truncated Lindley (LTL), double truncated Lindley (DTL) distributions are introduced.

The pdf of DTL distribution is denoted by \( g_D(x; \theta) \),

\[ g_D(x; \theta) = \frac{\theta^2 (1+x) e^{-\theta x}}{1 + \theta F(\zeta; \theta) - F(\zeta; \theta)}, \quad 0 \leq x \leq \zeta < \infty \]

When \( \theta < 1, g_D(x) \) is uni-modal and mode values is \( M_0 = (1-\theta) / \theta \)

The hazard rate function of UTL distribution is increasing in \( x \) and \( \theta \).

The pdf of LTL distribution is,

\[ g_L(x; \theta) = \frac{\theta^2 (1+x) e^{-\theta x}}{1 + \theta F(\zeta; \theta) - F(\zeta; \theta)}, \quad 0 \leq x \leq \zeta \]

The hazard rate function of UTL distribution is increasing in \( x \) and \( \theta \).

The truncated distributions can be quite effectively used to model the real problems and so we can use the truncated Lindley distributions in various fields including engineering, medical, finance and demography where such type of truncated data are commonly encountered. Among the three truncated versions DTL is more effective than UTL and LTL, also UTL is more effective than LTL.

Many researchers propose the truncated versions of the usual statistical distributions including, Ahmed et al.\(^{28}\) discussed the application of the truncated version of the Birnbaum–Saunders (BS) distribution to improve a forecasting actuarial model and particularly. Zaninetti\(^{28}\) discussed the application of the truncated Pareto distribution to the statistical analysis of masses of stars and of diameters of asteroids for modelling data from insurance payments that establish a
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**Discrete Lindley distribution**

Deniz proposed the discrete Lindley distribution (DLD). It is obtained by discretizing the continuous failure model of the Lindley distribution. The probability mass function (pmf) of DLD is

\[ f(x) = \frac{\lambda^x}{1 - \log \lambda} \left[ 2 \log \lambda + (1 - \lambda) \log \lambda^x \right] \quad x = 0, 1, \ldots \]

and the hazard rate function is

\[ h(x) = \frac{2 \log \lambda + (\lambda - 1) \log \lambda^x}{1 - (1 + x) \log \lambda} \]

The mean and variance of DLD is increase with \( \lambda \) and over-dispersed, therefore, more flexible than the Poisson distribution to model actuarial data that commonly include the over-dispersion phenomenon. It has an increasing hazard rate and it is unimodal. For large values of the parameter \( \lambda \), the mode moves to the right, showing a great versatility. Also if \( X \) follows a continuous Lindley distribution with parameter \( \theta > 0 \) then the random variable \( Y = \lfloor X \rfloor \) follows a DLD with parameter \( \lambda = e^{-\theta} \), here \( \lfloor . \rfloor \) denotes the integer part.

The researchers who were discretise various continuous distributions including.

**Poisson Lindley distribution**

A mixture distribution is the probability distribution of a random variable that is derived from a collection of other random variables. A compound distribution resembles in many ways the original distribution that generated it, but typically has greater variance, and often heavy tails as well. Any probability distribution on \([0, \infty)\) can function as the mixing distribution for a poisson mixture.

**Discrete poisson Lindley distribution**

The discrete Poisson Lindley distribution (DPLD) is the Poisson distribution compounded with Lindley distribution and is proposed by Sankaran. The pmf of DPLD is

\[ f(x) = P(X = x) = \frac{\theta^x (x + \theta + 2)}{(\theta + 1)^{x+2}} \quad x = 0, 1, \ldots, \theta > 0 \]

The DPLD is over dispersed, so it can be used in fields like biological science and medical science. In the field of genetics and ecology DPLD gives much closer fit than Poisson distribution and thus it can be considered as an important tool for modeling data in these field.

Ghitany discussed the estimation methods for the DPLD and its applications. The DPLD has been generalized by many researchers. Shanker studied two parameter Poisson–Lindley distribution with compounding Poisson distribution by mixing Poisson distribution with a two parameter Lindley distribution. Also Shanker studied Poisson–Lindley distribution and its application to biological science.

**Discrete poisson–Akash distribution (DPAD)**

Shanker proposed DPAD. The pdf of Poisson mixture of Akash distribution is

\[ f(x) = \frac{\theta^x x^2 + 3x + (\theta^2 + 2\theta + 3)}{(\theta + 1)^{x+3}} \quad x = 0, 1, 2, \ldots, \theta > 0 \]

DPAD has an increasing hazard rate and unimodal and always over-dispersed thus it is a suitable model for count data which are over-dispersed. DPAD gives much closer fit than Poisson distribution and DPLD in almost all cases, DPAD has some flexibility over DPLD.

Shanker et al. discussed Poisson–Akash Distribution and its Applications. Shanker has also introduced size based and zero truncated version of Poisson Akash distribution and studied their properties.

**Truncated lindley distribution**

Merovci proposed Transmuted Lindley Distribution. The pmf of Transmuted Lindley Distribution is,

\[ f(x) = \frac{\theta^x (1 + x)e^{-\theta x}}{(1 + \theta + x)e^{-\theta x}} \quad x = 0, \theta > 0 \]

The transmuted Lindley distribution is an extended model to analyze more complex data and it generalizes some of the widely used distributions, at \( \theta = 0 \) it gives Lindley distribution. The corresponding cdf is,

\[ F(x) = \frac{1 - \theta + x e^{-\theta x}}{(1 + \theta + x)e^{-\theta x}} \]

The hazard rate function is

\[ h(x) = \frac{\theta^x (1 + x)(1 + \theta + x)e^{-\theta x}}{(1 + \theta + x)^2} \]

From this if \( \lambda = 1 \), the hazard rate is decreasing. Aryal studied Transmuted Weibull distribution and its properties.

**Wrapped lindley distribution**

A wrapped probability distribution is a continuous probability distribution that describes data points that lie on a unit n–sphere. The cases of wrapped Lindley distribution have been studied extensively by Joshi. The wrapped Lindley (WL) random variable is defined as \( X \mod 2\pi \), such that for \( \theta \in [0, 2\pi] \), the pdf is given by

\[ g(\theta) = \frac{\lambda^2}{1 + \lambda} e^{-2\lambda \theta} \left[ \frac{1 + \theta}{1 - e^{-2\lambda \theta}} + \frac{2\pi e^{-2\lambda \theta}}{(1 - e^{-2\lambda \theta})^2} \right] \quad \theta \in [0, 2\pi], \lambda > 0 \]

The random variable \( \theta \) having wrapped Lindley distribution is denoted by \( \theta \sim WL(\lambda) \). The cdf of WL is

\[ G(\theta) = \frac{1}{1 - e^{-2\lambda \theta}} \left[ 1 - e^{-2\lambda \theta} - \frac{\lambda \theta}{1 + \lambda} \right] \frac{2\pi e^{-2\lambda \theta}}{1 - e^{-2\lambda \theta}} \theta \in [0, 2\pi], \lambda > 0 \]

Joshi showed that wrapped Lindley distribution give good fit to the data data set (orientations of 76 turtles after laying eggs and is given in Table 1 (Rao) than wrapped exponential distribution.

**Conclusion**

This paper studied the well–established and widely used Lindley distribution and its generalizations. The variety of generalizations
or parameters induction can be used to handle various real data sets with complex structure. These models will be useful for constructing probability models and may help the development of new classes from the Lindley distribution in future.

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**Conflict of interest**

The author declares that there in none of the conflicts.

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