Design and preliminary recruitment results of the Cluster randomised triAl of PSA testing for Prostate cancer (CAP)

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Background: Screening for prostate cancer continues to generate controversy because of concerns about over-diagnosis and unnecessary treatment. We describe the rationale, design and recruitment of the Cluster randomised triAl of PSA testing for Prostate cancer (CAP) trial, a UK-wide cluster randomised controlled trial investigating the effectiveness and cost-effectiveness of prostate-specific antigen (PSA) testing.

Methods: Seven hundred and eighty-five general practitioner (GP) practices in England and Wales were randomised to a population-based PSA testing or standard care and then approached for consent to participate. In the intervention arm, men aged 50–69 years were invited to undergo PSA testing, and those diagnosed with localised prostate cancer were invited into a treatment trial. Control arm practices undertook standard UK management. All men were flagged with the Health and Social Care Information Centre for deaths and cancer registrations. The primary outcome is prostate cancer mortality at a median 10-year-follow-up.

Results: Among randomised practices, 271 (68%) in the intervention arm (198 114 men) and 302 (78%) in the control arm (221 929 men) consented to participate, meeting pre-specified power requirements. There was little evidence of differences between trial arms in measured baseline characteristics of the consenting GP practices (or men within those practices).

Conclusions: The CAP trial successfully met its recruitment targets and will make an important contribution to international understanding of PSA-based prostate cancer screening.

Annually, over 500 000 men worldwide are diagnosed with prostate cancer, which is rapidly becoming the most common male cancer, due to an ageing population and increased detection through serum prostate-specific antigen (PSA) testing. However, the aetiology of prostate cancer is poorly understood and there are limited prospects for primary prevention (Frankel et al, 2003). Serum PSA testing followed by transrectal ultrasound-guided needle biopsy identifies cancers localised to the prostate gland, and

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could form the basis of a screening programme to detect disease amenable to cure. However, since many potentially identifiable prostate cancers are indolent (Sakr et al., 1996), over-diagnosis and over-treatment (Donovan, 2012; Sandhu and Andriole, 2012), which can impact adversely on a man’s quality of life, are major concerns (Neal et al., 2009).

Two recently published randomised controlled trials of PSA-based screening (Andriole et al., 2009; Schroder et al., 2009) failed to resolve the controversy over the effectiveness of screening in reducing mortality and improving quality of life (Neal et al., 2009). In the primary analysis of the European Randomised Study of Screening for Prostate Cancer (ERSPC), a 20% reduction in prostate cancer mortality was observed after a median follow-up of 9 years, with PSA testing every 4 years vs no screening (Schroder et al., 2009). The effect estimates were imprecise, however, with the 95% confidence limits indicating that the benefit could be as much as 35% or as little as 2%. The absolute reduction in risk suggested that 1410 men would need to be screened and 48 cancers detected to prevent one death from prostate cancer at approximately 8 years following diagnosis. A further analysis with median follow-up of 11 years (Schroder et al., 2012) yielded a similar 21% relative reduction in prostate cancer mortality and a slightly improved absolute benefit (1055 men would need to be screened and 37 cancers detected to prevent one death from prostate cancer). A microsimulation study based on ERSPC data suggested that 43% of cancers detected by screening are overdiagnosed and that the mortality benefits are diminished by a 23% reduction in prostate cancer mortality and improving quality of life (Neal et al., 2009).

The PLCO (prostate, lung, colon and ovarian cancer screening) trial reported no mortality benefit from screening after 7 (Andriole et al., 2009) or 13 years (Andriole et al., 2012) of follow-up. However, there are relatively few prostate cancer deaths reported to date. In addition, high levels of non-compliance in the intervention arm and contamination by widespread PSA testing in the control arm, make it unlikely that the PLCO trial can accurately estimate the magnitude of the intervention effect (Gulati et al., 2012).

A recently updated Cochrane review (Lluc et al., 2011) of five randomised controlled trials, a meta-analysis (Djulbegovic et al., 2010) of six randomised controlled trials, including the ERSPC, PLCO and 20-year follow-up of a small, long-term trial (Sandblom et al., 2011) do not provide strong evidence that screening causes reductions in all-cause or prostate cancer-specific mortality that are important enough to outweigh potential harms. The recent updated report from the US Preventive Services Task Force recommended against routine PSA-based screening (Chou et al., 2011). However, the controversy surrounding these draft recommendations (Brett and Ablin, 2011; McNaughton-Collins and Barry, 2011) suggests continuing concerns about the methodological quality of some evidence and associated risks of bias. These concerns are reinforced by favourable modelling projections (Gulati et al., 2011) and secondary analyses (Roobol et al., 2009); thus, the benefits and harms of population or targeted prostate cancer screening remain uncertain.

The most effective treatment for PSA-detected localised prostate cancer is also unknown because treatment within published screening trials was not randomised, and published treatment trials are difficult to interpret in relation to screening and the known heterogeneity of the disease. The findings from the SPCG-4 trial (comparing radical surgery and watchful waiting) provide robust evidence of a reduction in prostate cancer mortality in those assigned to radical surgery, but few men were recruited by PSA testing (Hofmager et al., 2012). The PIVOT trial did not provide conclusive evidence that older men with higher risk tumours (mean age 67 years, 48% Gleason 7 or higher, 66% intermediate/high risk) benefit from radical surgery compared with observation (Wilt, 2012). Neither trial has clear application for younger men with predominantly low-risk tumours in contemporary PSA-detected cohorts.

Here, we report the design, methods and results of recruitment in the CAP (Cluster randomised triAl of testing for Prostate cancer) trial. CAP is a nationwide, primary care-based cluster RCT that evaluates whether PSA testing of men aged 50–69 years will reduce prostate cancer mortality and be cost-effective (ISRCTN92187251). The embedded ProtecT trial (Donovan et al., 2003; Lane et al., 2010) (ISRCTN20141297) compares the effectiveness and cost-effectiveness of radical prostatectomy, radical conformal radiotherapy and active monitoring (regular PSA testing and treatment review) in men with PSA-detected clinically localised prostate cancer. The CAP trial will estimate the effect of a policy to invite men aged 50–69 years for PSA testing.

MATERIALS AND METHODS

Study design. CAP is a pragmatic cluster RCT, testing effectiveness in everyday practice by comparing an invitation to attend for population-based PSA testing for prostate cancer (the ProtecT trial intervention arm) with standard UK National Health Service (NHS) care (control arm) in men aged 50–69 years registered with general practices (community-based primary care clinics) in and around eight cities in England and Wales: Birmingham, Bristol, Cambridge, Cardiff, Leeds, Leicester, Newcastle and Sheffield (Figure 1).

Cluster randomisation. General practitioner (GP) practices (the cluster unit) were randomised between 2001 and 2007 by a statistician not otherwise involved in the study before their recruitment (‘Zelen’ design (Zelen, 1979)). Randomisation was blocked and stratified by geographical area based on groups of 10–12 neighbouring practices and using a computerised random number generator to allocate a near-equal number of practices in each stratum to intervention and comparison arms. Where a stratum contained an odd number of practices, the greater number of practices was randomised to the intervention arm to support recruitment to the embedded ProtecT trial treatment.

Recruitment of GP practices. Randomised practices were contacted by a researcher who briefed the GPs primary care doctors and practice managers, following which an information and consent pack was sent, tailored to the allocated arm. A senior partner (or equivalent) was asked to provide written consent for participation in the study, acting as a ‘guardian’ of the interests of individuals currently registered at the practice who would be involved in the study. Practices who contemporaneously recruited men to a trial of rofecoxib in the prevention of prostate cancer (Vioxx in Prostate Cancer Prevention study (www.clinicaltrials.gov/show/NCT00060476) were excluded. Control arm practices within clusters where no intervention arm practices were recruited,

![Figure 1. Trial design.](https://example.com/figure1.png)
and intervention arm practices in clusters where no control arm practices were recruited, were also excluded.

The success of randomisation and the impact of post-randomisation exclusions or dropouts were assessed by comparing routinely available practice level variables between the trial arms: list size; index of multiple deprivation (IMD) scores at lower level super output area (IMD England and Wales, 2011) based on practice postcode; urban–rural classification (Urban Rural Classification ONS, 2011); and mean prevalence of cancer, diabetes, obesity and coronary heart disease, based on data routinely submitted to the NHS Quality Outcomes Framework by GP practices (Quality Outcomes Framework (QOF) 2007/08, 2011).

Recruitment of participants. Recruitment to the intervention arm has been described in detail elsewhere (Donovan et al, 2002; Lane et al, 2010). Briefly, in each intervention arm practice, men aged 50–69 years were identified from the computerised patient lists on a specific date (Lane et al, 2010) (the ‘list date’). Each man was sent a single invitation by letter to attend a nurse-led prostate-check clinic appointment where the potential benefits and harms of having a PSA test were explained, and written informed consent was obtained. There was a second consent form completed after 24 h for all men who agreed to PSA testing. Men who attended the prostate-check clinic were also asked whether they would consent to being traced and flagged for vital status follow-up with the Health and Social Care Information Centre (HSCIC).

In control arm practices, all men aged 50–69 years registered with each practice on a pre-specified date (consistent with the list dates in intervention arm practices within the same stratum) were eligible for inclusion. Men were not individually invited to participate as they were undergoing standard NHS management. However, they (and intervention arm non-responders) were given the opportunity to opt out of follow-up with the HSCIC through information provided at each practice.

Lists of all men aged 50–69 years who were registered at the intervention and control arm practices on the ‘list dates’, and who had not opted out, were obtained through each practice computer system and the following data were extracted: name, postcode, date of birth, NHS number and practice identification number. These data were encrypted and returned to the co-ordinating centre (School of Social and Community Medicine, University of Bristol). These lists constitute the CAP trial cohort for follow-up.

Trial interventions. In the intervention arm, men with a raised PSA (≥3.0 ng ml$^{-1}$) were invited for diagnostic tests including a transrectal ultrasound (TRUS)-guided biopsy. All men subsequently diagnosed with clinically localised prostate cancer (T1-T2, NX, M0) were asked to consider randomisation within the ProtecT three-arm trial of treatments: radical prostatectomy, radical conformal radiotherapy and active monitoring (regular PSA testing and review). If randomisation was unacceptable, a patient preference treatment option was agreed (Donovan et al, 2002; Lane et al, 2010). Men found to have advanced disease were treated in the NHS urology clinics of the recruiting centres. All men treated via randomisation or by preference were followed up in annual nurse-led clinics.

Control arm practices were supplied with contemporary information from the NHS prostate cancer-risk management programme (NHS, 2011) to encourage standard NHS management, including the provision of standardised information about the potential benefits and harms of PSA testing for any man over the age of 50 years who requested a PSA test.

Follow-up and identification of a prostate cancer-related event. Surveillance for relevant outcomes is passive, triggered by the occurrence of death or cancer registration identified via flagging with the HSCIC. When such triggers are received, clinical and resource-use data from the man’s medical records are obtained, blind to the cancer registry or death certificate information.

Data items include: symptoms and signs of prostate cancer presence and progression; diagnostic and monitoring tests; histology, tumour stage; treatments received and the outcomes of treatment; complications related to prostate cancer diagnosis or treatment; co-morbidities; and other resource-use data (e.g., length of inpatient stay, outpatient appointments). For men who have died, a researcher, blind to information on the death certificate, prepares a short-structured vignette that summarises the key clinical events from diagnosis to death.

Determination of cause of death. To minimise potential attribution bias (Albertsen et al, 2000; Black et al, 2002), an international Cause of Death Evaluation (CODE) Committee reviews clinical information abstracted from medical records (vignettes) on all possible prostate cancer deaths in both arms of the trial (independent chair: professor Peter Albertsen, Connecticut, USA). A list of specific ICD9 and ICD10 codes, adapted from those used by the PLCO Screening Trial (Miller et al, 2000), is used to identify a possible prostate cancer event from the ICD codes listed in parts I and II of each man’s death certificate. The following outcomes are assigned by the reviewers (blind to trial arm and information on the death certificate), according to pre-defined definitions adapted from those used in the ERSPC trial (de Koning et al, 2003): definite, probable, possible, unlikely or definitely not prostate cancer; and definite or probable intervention-related mortality. Further information about the CODE process and algorithm will be published separately.

Primary and secondary outcome measures. The primary outcome is ‘definite’ or ‘probable’ prostate cancer mortality after a median 10-year-follow-up, which will occur in 2016. Secondary outcomes include all-cause mortality at median follow-up of 10 and 15 years; definite or probable prostate cancer mortality at 15 years; disease status and staging; and cost-effectiveness. Sensitivity analyses will restrict prostate cancer mortality to deaths classified as definite prostate cancer and widen the definition to include definite, probable and possible prostate cancer deaths. The impact of PSA testing on health-related quality of life has been investigated in a linked study (Avery et al, 2008a,b; Macefield et al, 2010).

Assessment of contamination. The level of PSA testing from a convenient sample of 87 control arm general practices in 2007 was collected using the EMIS LV computer system. Information for a total of 126,716 men aged 45–89 years with no recorded diagnosis of prostate cancer before 1 January 2007 was collected and has been reported previously (Williams et al, 2011). Further studies of contamination are underway.

Ethics and research governance. Men who attended the prostate-check clinic in the intervention arm gave individual informed consent (Trent MREC/01/4/025). Approval for flagging of men in the control arm and non-responders in the intervention arm was obtained under Section 251 of the NHS Act 2006 (PIAG 4-09 (k)/2003) from the UK Patient Information Advisory Group (PIAG) (which became until recently the National Information Governance Board Ethics and Confidentiality Committee, NIGBEC)). Trent MREC provided ethics approval for flagging (MREC/03/4/093) and review of the medical records of men with prostate cancer (05/MRe04/78). Men who were alive when we were notified of a prostate cancer diagnosis provided individual informed consent for review of their medical records. PIAG/NIGBEC approval allows review of the medical records of men who died of a cause potentially related to prostate cancer before consent could be gained (provided the man did not record an objection to their medical records being used for research while alive) (PIAG 1-05(0)/2006). All clinical centres have local research governance and ethics approval. The CAP study is sponsored by the University of Bristol and funded by Cancer Research UK and the Department of Health.
Analysis plan. The primary analysis will be by intention-to-screen: all men followed up in the intervention arm will be included in that arm for analysis regardless of whether or not they responded to the invitation to have a PSA test. Similarly, control arm men will be included in that arm regardless of whether they had a PSA test. The effect of an invitation to screening on prostate cancer mortality will be estimated as a rate ratio while allowing for clustering by general practice once a median of 10 years follow-up data has accumulated. Analyses will be adjusted for randomisation strata.

Follow-up will commence on the ‘list date’ for each practice. Men will be excluded from the analyses if they had prevalent prostate cancer (diagnosed at or before the list date). Men will be censored on the date of death, or date of loss to follow-up (e.g., notified by the HSCIC as having emigrated). This approach to analysis will be adapted to the secondary outcome measure, such as all-cause mortality. Sensitivity analyses will assess whether controlling for any baseline imbalances in practice characteristics makes any difference to the estimated effects of screening.

If any of the three treatment arms in the ProtecT trial is superior (leads to lower mortality compared with the other arms), differences in prostate cancer or all-cause mortality between the CAP trial arms may be smaller than would be expected if a screening programme had used the optimal treatment(s). We will estimate the effect on mortality of such an ‘optimal’ screening programme, based on the (unbiased) treatment effect estimates from ProtecT and the (unbiased) overall effect estimates from CAP. Similar to what has been done for the ERSPC study (Roobol et al, 2009), we will use appropriate methods to estimate the effect of testing in those who undergo PSA measurement (complier average causal effect). This can be used to predict the overall effect of a screening programme under different assumptions about uptake of PSA testing and subsequent treatment.

Sample size and power. The sample size of the CAP trial was determined by the number of practices required to meet the recruitment target for the embedded ProtecT trial. Power calculations assessed whether this would be sufficient to detect reductions in prostate cancer mortality of a magnitude likely to be important to the NHS (Gohagan et al, 1994; Auvinen et al, 1996). We considered a range of plausible between-practice coefficients of variation (a measure of the amount of between-practice variation in prostate cancer mortality rates). For example, based on estimates of the incidence of prostate cancer mortality (the primary outcome) over 10 years using 2006 England and Wales incidence and mortality data (Office for National Statistics (ONS), 2008), and assuming a plausible coefficients of variation of 0.2, 209 000 men in each study arm allows a true overall prostate cancer mortality rate ratio of 0.87 to be detected with 80% power at 5% significance. This corresponds to an intervention rate ratio among men undergoing screening (IRR) from 0.62 to 0.73, assuming uptake of the invitation to PSA testing of 35% and 50%, respectively. These IRRs are similar to those assumed in the power calculations for the ERSPC trial (de Koning et al, 2002), where post-consent randomisation ensured most men would undergo testing. Our estimates of the impact of a range of PSA contamination rates (defined as ever having a PSA test during follow-up, as in ERSPC (de Koning et al, 2002)), suggest that the effect of contamination is minimal unless it reaches 20%, a level unlikely in this trial (Williams et al, 2011).

RESULTS

Recruitment

GP practice recruitment. In total, 911 GP practices were randomised in 99 geographical areas. Six areas (52 practices) were excluded from further analysis as no intervention or control arm practices were recruited (see Figure 2). Of the 440 practices in the intervention arm and 419 in the control arm, 42 and 32 (respectively) were ineligible, because the GP practice was involved in the Vioxx trial, had ceased to exist, consented too late to allow participation, exclusively covered a atypical population (e.g., private, asylum seekers, children in care, homeless, sheltered housing), could not produce a list of men aged 50–69, or had been randomised in error (i.e., practices that participated in the feasibility phase of the ProtecT trial (Donovan et al, 2003)). Of the remaining 398 intervention arm and 387 control practices, 42 and 40, respectively, explicitly refused to participate, while 85 and 45, respectively had not responded after being allowed at least one year to consider taking part (and can be considered as implicit refusals, generated by the way the study was carried out). There was little evidence of differences between measured characteristics of the 573 participating practices and the 130 implicit refusal practices (list size 6827 vs 7264 (P = 0.4); IMD English 29 vs 27 (P = 0.3); IMD Wales 21 vs 23 (P = 0.8); number of urban practices 88% vs 82% (P = 0.07); cancer prevalence 0.55% vs 0.54% 9 (P = 0.7); diabetes prevalence 3.7% vs 3.6% (P = 0.7); obesity prevalence 7.9% vs 8.5% (P = 0.08); coronary heart disease 4% vs 3.8% (P = 0.2), respectively). Where practices explicitly refused, the major reasons were lack of interest, time or space (Down et al, 2009). Thus, among the eligible practices, 271 (68%) in the intervention arm and 302 (78%) in the control arm consented to participate.

We found little evidence of differences between the measured characteristics between the 785 eligible practices randomised to the intervention compared with control arms. Neither did we find evidence of differences in measured characteristics between the practices who consented to participate in the intervention compared with control arms (Table 1).

Take-up of PSA testing. Among 197 925 eligible men in the intervention arm, 16 752 were not invited for a PSA test because they met ProtecT study exclusion criteria (they were excluded by their GP because of terminal or severe mental illness, serious co-morbidity or pre-existing prostate cancer diagnosis). However, these men are included in the follow-up for prostate cancer incidence and mortality via the HSCIC. A further 1 404 men were excluded as they had prevalent prostate cancer and thus could not trace 47 men, these figures are subject to change as a result of continued updates from the HSCIC. Of the remaining 179 722 men who received an invitation to PSA testing within the intervention arm, 79 889 (45%) attended a prostate-check clinic and 65 113 (36%) men gave full consent for their PSA level to be measured as part of the ProtecT study.

In the control arm among 221 629 men, three men opted out, 1 632 had prevalent prostate cancer and HSCIC failed to trace 81, as above these figures are subject to change upon receipt of further information from the HSCIC. Thus, a total of 416 387 men were randomised into the intervention and control arms (Figure 2), a population-based sample covering approximately 8% of the England and Wales male population aged 50–69 years (Quality Outcomes Framework (QOF) 2007/08, 2011). Consent rates for flagging with the HSCIC were over 99%, and 97.5% of men were matched to the NHSCR database automatically with 2.5% requiring manual matching. There were no substantial differences between the men in the intervention and control arms in terms of age (mean 59.1), baseline prevalence of prostate cancer (prostate cancer diagnosed before randomisation was 0.71% in the intervention arm compared with 0.74% in the control arm) or index of multiple deprivation score, a census-derived measure of multiple aspects of area-level deprivation, in England (22.9) or Wales (20.8 vs 19.6, respectively).
### DISCUSSION

The CAP trial recruited successfully over 415,000 men from 573 primary care practices across the UK. Its future contribution to the evidence base is illustrated by the ways in which it differs in methods of recruitment, screening and treatments offered compared with the two major published screening trials (summarised in Table 2; Schroder and Roobol, 2010). Cluster randomisation was used because PSA testing was offered involving screening: 13 Ceased to exist: 19

Cluster randomisation was used because PSA testing was offered to all men in intervention arm practices where individual randomisation would have been impracticable. Compared with an individually randomised design, the clustered design reduces the likelihood of contamination in the control group: only 6.2% of men (95% confidence interval: 5.7–7.0%; practice-based interquartile range 3.6–8.4%) had their PSA tested in 2007 in a sample of adults aged 50–64 years enrolled in the ProtecT trial. Consent rates amongst intervention and control arm men were traced and flagged with the HSCIC, ensuring an unbiased intention-to-screen analysis, comparing outcomes for the two arms of the trial according to the original random allocation.

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Figure 2. Consort diagram for recruitment into the Cluster Randomised Trial of Testing for Prostate Cancer (CAP), England and Wales.

| Practices randomised to intervention arm | n = 440 (93 areas) |
|-----------------------------------------|-------------------|
| Practices eligible for the intervention arm | n = 398 (93 areas) |
| Men participating in intervention arm | n = 196,474c |
| Men excluded from intervention* | Prostate cancer pre-randomisation n = 1,404 (0.71%) |
| Refused: 127 (32%) | Failed to trace at NHS IC n = 47 |
| Explicit refusal: 42 | Implicit refusal: 85 |

| Practices randomised to control arm | n = 419 (93 areas) |
|-------------------------------------|-------------------|
| Practices eligible for the control arm | n = 387 (93 areas) |
| Men participating in control arm | n = 219,913 |
| Men excluded from control* | Prostate cancer pre-randomisation n = 1,632 (0.74%) |
| Refused: 85 (22%) | Failed to trace at NHS IC n = 81 |
| Explicit refusal: 40 | Implicit refusal: 45 |

| n = 911 GP practices randomised (within 99 geographical areas) |
|----------------------------------------------------------------|
| Areas excluded because of zero cell within cluster: |
| Not approached, recruiting centre closed: 1 area (10 practices) |
| No consented intervention arm practices: 2 areas (18 practices) |
| No consented control arm practices: 3 areas (24 practices) |

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Cluster randomisation was used because PSA testing was offered to all men in intervention arm practices where individual randomisation would have been impracticable. Compared with an individually randomised design, the clustered design reduces the likelihood of contamination in the control group: only 6.2% of men (95% confidence interval: 5.7–7.0%; practice-based interquartile range 3.6–8.4%) had their PSA tested in 2007 in a sample of adults aged 50–64 years enrolled in the ProtecT trial. Consent rates amongst intervention and control arm men were traced and flagged with the HSCIC, ensuring an unbiased intention-to-screen analysis, comparing outcomes for the two arms of the trial according to the original random allocation.
The CAP trial aims to evaluate the impact of population-based PSA testing for prostate cancer—a test with acknowledged limited specificity (Holmstrom et al., 2009), imperfect sensitivity (15.2% of men with a PSA level of 4.0 ng ml\(^{-1}\) or less biopsied at the end of follow-up in the Prostate Cancer Prevention Trial were diagnosed with prostate cancer, with high-grade cancers (at least Gleason 7) observed throughout this range of PSA values) (Thompson et al., 2004), and a high rate of over-diagnosis of low-risk disease resulting in over-treatment. Currently, however, there are no alternative reliable screening methods. The modest PSA testing rate in the intervention arm reflects that achievable with a single invitation to eligible participants, but randomisation allows us to employ appropriate statistical methods (e.g., complier average causal effects, contamination-adjusted ITT (Sussman and Hayward, 2010)) to estimate the effect in those who undergo PSA testing (Cuzick et al., 1997; Roobol et al., 2009). This can be used to predict the overall effect of a screening programme under different assumptions about PSA uptake. We will not be able to measure quality of life in the men being followed up passively; however, a linked study has investigated the impact of PSA testing on health-related quality of life in men followed up in the intervention arm (Avery et al., 2008a,b; Macefield et al., 2010).

ERSPC (Schroder et al., 2009) and PLCO (Andriole et al., 2009) included various screening intervals, ranging from 2 to 4 years. Cancers identified in second and subsequent screening rounds were randomised (population-based effectiveness trial) in Sweden and Italy, all men identified from cancer registries consent underwent randomisation (efficacy trial). In Finland, Switzerland, Spain, only men giving informed consent underwent randomisation (efficacy trial). In Netherlands, Belgium, only men giving informed consent were randomised (efficacy trial). In England, Sweden and Italy, all men identified from cancer registries were randomised (population-based effectiveness trial). In Finland, Italy, all men identified from cancer registries were randomised (population-based effectiveness trial). In Netherlands, Belgium, Switzerland, Spain, only men giving informed consent were randomised (efficacy trial). In Finland, Sweden and Italy, all men identified from cancer registries were randomised (population-based effectiveness trial). In Netherlands, Belgium, Switzerland, Spain, only men giving informed consent were randomised (efficacy trial).

### Table 1. Comparison of baseline characteristics between participating intervention and control arm practices, England and Wales

|                      | Intervention arm (n = 271) | Control arm (n = 302) | P-value for difference* |
|----------------------|---------------------------|-----------------------|-------------------------|
| Mean practice list size (s.d.) | 6883 (3442)               | 6777 (3707)           | 0.7                     |
| Number of urban practices (s.d.) | 267 (88%)                 | 267 (88%)             | 0.8                     |

### Indices of deprivation

|                      | Intervention arm (n = 271) | Control arm (n = 302) | P-value for difference* |
|----------------------|---------------------------|-----------------------|-------------------------|
| Mean IMD score in England (s.d.) | 28 (19)                   | 30 (19)               | 0.2                     |
| Mean IMD score in Wales (s.d.) | 21 (14)                   | 22 (16)               | 0.7                     |

### Mean prevalence from QOF

|                      | Intervention arm (n = 271) | Control arm (n = 302) | P-value for difference* |
|----------------------|---------------------------|-----------------------|-------------------------|
| All cancers (s.e.)   | 0.57% (0.02%)             | 0.53% (0.01%)         | 0.07                    |
| Diabetes (s.e.)      | 3.6% (0.07%)              | 3.7% (0.06%)          | 0.4                     |
| Obesity (s.e.)       | 8.0% (0.2%)               | 7.8% (0.2%)           | 0.3                     |
| Coronary heart disease (s.e.) | 4.1% (0.09%)             | 3.9% (0.08%)         | 0.2                     |

Abbreviations: IMD = index of multiple deprivation; QOF = quality outcomes framework; s.d. = standard deviation; s.e. = standard error.

*P-values were calculated using Robust Regression, all analyses adjusted for geographic and primary care grouping.

**Rural/urban classification 2004 (scale: 1: urban; 2: town and fringe; 3: village; 4: hamlet and isolated dwellings).

*Prevalence of cancer from quality outcomes framework 2004/5 where available (n = 491).
tended to be small with favourable histological grading, and with a low probability (about 1 in 10) of becoming clinically important in the man’s lifetime (Schroder et al, 2010). The CAP trial aims to assess the impact of single PSA testing, allowing the lifetime risk of developing clinically important prostate cancer amongst men initially screen negative to be quantified, and contributing natural history data to the debate on whether a single PSA measurement in middle age could be used to risk stratify men for further PSA testing and targeted treatment (Vickers et al, 2010). It is possible that a single screen may underestimate the potential long-term benefits of a PSA-testing programme with multiple rounds (Hanley, 2005), as investigated in the ERSPC and other trials, and might be considered difficult to implement in countries with established PSA-testing programmes. A single PSA test, however, is less likely to identify large numbers of clinically insignificant disease (Heijnsdijk et al, 2012). In the UK, a previously unscreened population, the pool of clinically significant disease will be larger than in a population already undergoing regular screening. The trial design, therefore, will determine the effectiveness of a single PSA test while minimising the negative public health consequences of identifying clinically insignificant disease.

Misattribution of cause of death (Albertsen et al, 2000; Black et al, 2002) is an important potential source of bias in trials with cancer-specific mortality as the primary end point. Small inaccuracies in attributing cause of death can have the potential to alter the results if misclassification is differential between study arms. To minimise misclassification of cause of death, possible prostate cancer deaths are reviewed by an independent international cause of death committee blinded to the trial arms, with an embedded quality assurance process to ensure the reliability of the vignettes and the decisions of the reviewers (data to be presented separately).

We are powered to detect an overall prostate cancer mortality rate ratio of 0.87 (13% relative reduction). Reductions in prostate cancer mortality of the order of 15–20% are likely to be important to the NHS. Although our estimates suggest the effect of contamination is minimal unless it reaches 20%, the power of the trial will be reduced if men in the control practices are screened for prostate cancer (‘contamination’). An advantage of the cluster randomised design is that contamination is likely to be less of a problem than would be the case if men were individually randomised, and hence were alerted to the possibility of being screened for prostate cancer.

The CAP and ProtecT trials have been designed and conducted to make a much awaited and crucial contribution in understanding the effectiveness and cost-effectiveness of prostate cancer screening, as well as directly informing international policy on testing and treatment of this common malignancy.

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REFERENCES

Albertsen PC, Walters S, Hanley JA (2000) A comparison of cause of death determination in men previously diagnosed with prostate cancer who died in 1985 or 1995. J Urol 163: 519–523.

Andriole GL, Crawford ED, Grubb III RL, Buys SS, Chia D, Church TR, Fouad MN, Gelman EP, Kvale PA, Reding DJ, Weissfeld JL, Yokochi LA, O’Brien B, Clapp JD, Rathmell JM, Riley TL, Hayes RB, Kramer BS, Izmirlian G, Miller AB, Pinsky PF, Prorok PC, Gohagan JK, Berg CD (2009) Mortality results from a randomized prostate-cancer screening trial. N Engl J Med 360: 310–319.

Andriole GL, Crawford ED, Grubb III RL, Buys SS, Chia D, Church TR, Fouad MN, Isaacs C, Kvale PA, Reding DJ, Weissfeld JL, Yokochi LA, O’Brien B, Ragard LR, Clapp JD, Rathmell JM, Riley TL, Hsing AW, Izmirlian G, Pinsky PF, Kramer BS, Miller AB, Gohagan JK, Prorok PC (2012) Prostate cancer screening in the randomized Prostate, Lung, Colorectal, and Ovarian Cancer Screening Trial: mortality results after 13 years of follow-up. J Natl Cancer Inst 104: 125–132. Auvine A, Riebergen JB, Deniz LJ, Schroder FH, Prorok PC (1996) Prospective evaluation plan for randomised trials of prostate cancer screening. The International Prostate Cancer Screening Trial Evaluation Group. J Med Screen 3: 97–104.

Avery KN, Blazey JM, Lane JA, Neal DE, Hamdy FC, Donovan JL (2008a) Decision-making about PSA testing and prostate biopsies: a qualitative study embedded in a primary care randomised trial. Eur Urol 53: 1186–1193.

Avery KN, Metcalfe C, Blazey JM, Lane JA, Neal DE, Hamdy FC, Donovan JL (2008b) Prostate-specific antigen testing and prostate biopsy: are self-reported lower urinary tract symptoms and health-related quality of life associated with the decision to undergo these investigations? BJU Int 102: 1629–1633.

Black WC, Haggstrom DA, Welch HG (2002) All-cause mortality in randomized trials of cancer screening. J Natl Cancer Inst 94: 167–173.

Brett AS, Ablin RJ (2011) Prostate-cancer screening—What the U.S. Preventive Services Task Force left out. N Engl J Med 365: 1949–1951.

Chou R, CrossRef JM, Dana T, Bougatsos C, Blazina I, Fu R, Gleisner K, Koenig HC, Lam C, Malta A, Ruggie JB, Lin K (2011) Screening for prostate cancer: a review of the evidence for the U.S. Preventive Services Task Force. Ann Intern Med 155: 762–771.

Cuzick J, Edwards R, Segnan N (1997) Adjusting for non-compliance and contamination in randomized clinical trials. Stat Med 16: 1017–1029.

de Koning HJ, Blom J, Merkelbach JW, Raaijmakers R, Verhaegen H, Van VP, Nelen V, Coebergh JW, Hermans A, Ciatto S, Makinen T (2003) Determining the cause of death in randomized screening trial(s) for prostate cancer. BJU Int 92(Suppl 2): 71–78.

de Koning HJ, Lien T, Haan CA, Boer R, Schroder FH, Alexander FE. ERSPC (2002) Prostate cancer mortality reduction by screening: power and time frame with complete enrollment in the European Randomised Screening for Prostate Cancer (ERSPC) trial. Int J Cancer 98: 268–273.

Djulbegovic M, Beyth RJ, Neuberger MM, Stoffis TL, Vieweg J, Djulbegovic B, Dahm P (2010) Screening for prostate cancer: systematic review and meta-analysis of randomised controlled trials. BMJ 341: c5453.

Donovan JL, Hamdy FC, Neal DE, Peters TJ, Oliver S, Brindle L, Jewell D, Powell P, Gillatt D, Dedman D, Mills N, Smith M, Noble S, Lane A (2003) Prostate testing for cancer and Treatment (ProtecT) feasibility study. Health Technol Assess 7: 1–42.

Donovan J, Mills N, Smith M, Brindle L, Jacoby A, Peters T, Frankel S, Neal D, Hamdy F (2002) Quality improvement report: Improving design and
conduct of randomised trials by embedding them in qualitative research: ProtecT (prostate testing for cancer and treatment) study. Commentary: presenting unbiased information to patients can be difficult. BMJ 325: 766–770.

Donovan JL (2012) Presenting treatment options to men with clinically localized prostate cancer: the acceptability of active surveillance/monitoring. J Natl Cancer Inst Monogr 2012: 191–196.

Down L, Metcalfe C, Avery K, Noble S, Lane JA, Neal DE, Hamdy FC, Donovan JL (2009) Factors distinguishing general practitioners who more readily participated in a large randomized trial were identified. J Clin Epidemiol 62: 67–73.

Frankel S, Davey Smith G, Donovan J, Neal D (2003) Screening for prostate cancer. Lancet 361: 1122–1128.

Gohagan JK, Prorok PC, Kramer BS, Cornett JE (1994) Prostate cancer screening in the prostate, lung, colorectal and ovarian cancer screening trial of the National Cancer Institute. J Urol 152: 1905–1909.

Gulati R, Mariotto AB, Chen S, Gore JL, Etzioni R (2011) Long-term projections of the harm-benefit trade-off in prostate cancer screening are more favorable than previous short-term estimates. J Clin Epidemiol 64: 1412–1417.

Gulati R, Tsodikov A, Wever EM, Mariotto AB, Heijnsdijk EA, Katcher J, de Koning HJ, Etzioni R (2012) The impact of PLCO control arm contamination on perceived PSA screening efficacy. Cancer Causes Control 23: 827–835.

Hanley JA (2005) Analysis of mortality data from cancer screening studies: looking in the right window. Epidemiology 16: 786–790.

Heijnsdijk EA, Weer EM, Auvinen A, Hugosson J, Ciato S, Nelen V, Kwiatkowski M, Villers A, Paer A, Moss SM, Zappa M, Tammla TL, Makinen T, Carlsson S, Korfage II, Essink-Bot ML, Otto SJ, Draisma G, Bangma CH, Roobol MJ, Schroder FH, de Koning HJ (2012) Quality-of-life effects of prostate-specific antigen screening. N Engl J Med 365: 955–965.

Holmberg L, Bill-Axelson A, Steineck G, Garmo H, Palmgren J, Johansson E, Adami HO, Johansson JE (2012) Results from the Scandinavian prostate cancer group trial number 4: a randomized controlled trial of radical prostatectomy versus watchful waiting. J Natl Cancer Inst Monogr 2012: 230–233.

Holmstrom B, Johansson M, Bergh A, Stennan UH, Hallmans G, Stattin P (2009) Prostate specific antigen for early detection of prostate cancer: longitudinal study. BMJ 339: b3537.

Ilic D, O’Connor D, Green S, Wilt TJ (2011) Screening for prostate cancer: an updated Cochrane systematic review. BJU Int 107: 882–891.

IMD England and Wales (2011) http://neighbourhood.statistics.gov.uk or http://wales.gov.uk/topics/statistics/theme/wind.

Lane JA, Hamdy FC, Martin RM, Turner EL, Neal DE, Donovan JL (2010) Latest results from the UK trials evaluating prostate cancer screening and treatment: The CAP and ProtecT studies. Eur J Cancer 46: 3095–3019.

Macefield RC, Metcalfe C, Lane JA, Donovan JL, Avery KN, Blazey JM, Down L, Neal DE, Hamdy FC, Vickers A (2010) Impact of prostate cancer testing: an evaluation of the emotional consequences of a negative biopsy result. Br J Cancer 102: 1335–1340.

McNaughton-Collins MF, Barry MJ (2011) One Man at a Time–Resolving the contamination on perceived PSA screening efficacy. BMJ 339: b3537.

Minasian LM, Ford LG, Lippman SM, Crawford ED, Crowley JJ, Colman JR (2004) Prevalence of Prostate Cancer among Men with a Prostate-Specific Antigen Level ≤ 4.0ng per Millilitre. N Engl J Med 350: 2239–2246.

Urban rural classification ONS (2011) http://neighbourhood.statistics.gov.uk or http://www.statswales.wales.gov.uk/.

Vickers AJ, Cronin AM, Bjork T, Manor J, Nilsson PM, Dahlin A, Bjartell A, Scardino PT, Ulmert D, Lilja H (2010) Prostate specific antigen concentration at age 60 and death or metastasis from prostate cancer: case-control study. BMJ 340: c2073.

Williams N, Hughes JI, Turner EL, Donovan JL, Hamdy FC, Neal DE, Martin RM, Metcalfe C (2011) Prostate-specific antigen testing rates remain low in UK general practice: a cross-sectional study in six English cities. BJU Int 108: 1402–1409.

Wilt TJ (2012) The Prostate Cancer Intervention Versus Observation Trial:VA/NCI/AHRQ Cooperative Studies Program #407 (PIVOT): Design and Baseline Results of a Randomized Controlled Trial Comparing Radical Prostatectomy With Watchful Waiting for Men With Clinically Localized Prostate Cancer. J Natl Cancer Inst Monogr 2012: 184–190.

Zelen M (1979) A new design for randomized controlled trials. N Engl J Med 300: 1242–1245.

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