On the Origin of Switchbacks Observed in the Solar Wind

F. S. Mozer\textsuperscript{1,2}, S. D. Bale\textsuperscript{1,2}, J. W. Bonnell\textsuperscript{1}, J. F. Drake\textsuperscript{3}, E. L. M. Hanson\textsuperscript{1}, and M. C. Mozer\textsuperscript{4}

\textsuperscript{1}Space Sciences Laboratory, University of California, Berkeley, CA, USA
\textsuperscript{2}Physics Department, University of California, Berkeley, CA, USA
\textsuperscript{3}University of Maryland, College Park, MD, USA
\textsuperscript{4}Google Research, Brain Team, Mountain View, CA, USA

Abstract

The origin of switchbacks in the solar wind is discussed in two classes of theory that differ in the location of the source being either near the transition region near the Sun or in the solar wind itself. The two classes of theory differ in their predictions of the switchback rate (the number of switchbacks observed per hour) as a function of distance from the Sun. To distinguish between these theories, one-hour averages of Parker Solar Probe data were averaged over five orbits to find the following: (1) The hourly averaged switchback rate was independent of distance from the Sun. (2) The average switchback rate increased with solar wind speed. (3) The switchback size perpendicular to the flow increased as \( R \), the distance from the Sun, while the radial size increased as \( R^2 \), resulting in an increasing switchback aspect ratio with distance from the Sun. (4) The hourly averaged and maximum switchback rotation angles did not depend on the solar wind speed or distance from the Sun. These results are consistent with switchback formation in the transition region because their increase of tangential size with radius compensates for the radial falloff of their equatorial density to produce switchback rates that are independent of radial distance. This constant switchback rate is inconsistent with an in situ source. The switchback size and aspect ratio, but not their hourly average or maximum rotation angle, increased with radial distance to 100 solar radii. Additionally, quiet intervals between switchback patches occurred at the lowest solar wind speeds.

1. Introduction

Switchbacks are rotations of the magnetic field in the solar wind (Yamauchi et al. 2004; Landi et al. 2006; Suess 2007; Matteini et al. 2014; Borovsky 2016). They have been observed in abundance on the Parker Solar Probe, and they have been described extensively (Bale et al. 2019; Kasper et al. 2019; Dudok de Wit et al. 2020; Horbury et al. 2020; Krasnoselskiikh et al. 2020; McManus et al. 2020; Mozer et al. 2020; Tenerani et al. 2020).

Two classes of theory attempt to explain the origin of switchbacks. In one class of theory, processes in the transition region, such as magnetic field reconnection, field line stirring, nanoflares, etc., eject switchbacks that propagate into the solar wind (Axford & McKenzie 1997; Fisk 2005; De Pontieu et al. 2009; Samanta et al. 2019; Drake et al. 2021; Fisk & Kasper 2020; Magyar et al. 2021a, 2021b). In the other class of theory, processes in the solar wind itself, such as strong turbulence, excitation of the Kelvin–Helmholtz instability, or magnetic field expansion, produce the switchbacks (Malagoli et al. 1996; Baty & Keppens et al. 2002; DeForest et al. 2016, 2018; Chhiber et al. 2018; Ruffolo et al. 2020; Squire et al. 2020). Because the radial dependence of the switchback rate differs in the two classes of theory, this dependence was investigated in order to distinguish between the transition region and in situ models for the switchback origin.

2. Definition of a Switchback

Figure 1 presents three components of the magnetic field during 25 days near perihelion three, as measured in the spiral coordinate system. In this coordinate system, which is used for all data in this paper, the \( X \)-direction is perpendicular to the Parker spiral in the ecliptic plane and points in the direction of solar rotation (against the ram direction), \( Y \) is perpendicular to the ecliptic plane and points southward, and \( Z \) points inward along the Parker spiral. The Parker spiral for this data was computed for a solar wind speed of 400 km s\(^{-1}\). The spiky structures in Figure 1 indicate the possible locations of the switchbacks, as defined below, whose radial occurrence frequency will be determined.

The term “switchbacks” has been defined in a variety of ways such that the number of switchbacks in a given time interval can vary by more than an order of magnitude, depending on the definition. This variation is unimportant for the present purposes because the switchback rate, and not the number of switchbacks, is the topic of discussion. Switchbacks are defined in this paper as rotations of the magnetic field from the Parker spiral through more than 90°. They must also have durations greater than 6 s in order to eliminate signals from dust, time-domain structures, telemetry noise, low-frequency waves, etc. To complete the definition, it is necessary to define the cutoff angle such that, when the magnetic field rotation decreases below this angle, the switchback has ended. In Figure 2, the number of switchbacks in the data of Figure 1 are plotted as a function of this cutoff angle. For a cutoff angle near 90°, unwanted, small excursions of substructures of field rotation above and below 90° are captured. For a cutoff angle near 15°, there are too few switchbacks for statistical significance. A switchback cutoff angle of 60° was used to analyze the switchbacks in this paper. (A cutoff angle of 75° was tested on subsets of the data with the result that the switchback numbers were greater for a 75° cutoff than for a 60° cutoff, see Figure 2, but the slopes versus other quantities were the same as those described below for the 60° cutoff.)
3. Data Analysis

A 2 minute interval of magnetic field rotation angles measured during an extremely active period is presented in Figure 3. The average number of switchbacks per hour will be shown to be about two and this interval has four switchbacks in 2 minutes. The black circles at the top of the figure identify these switchback start times. For a cutoff angle near 90°, the number of switchbacks can be estimated by eye to be about 12, while, for a cutoff angle of 15°, there is only one switchback during this time interval. Switchbacks may also be defined as abrupt changes of the magnetic field direction. According to this definition, examination by eye suggests that there may be four or five switchbacks in the interval shown.

Because the switchback rate may be a function of both the solar wind speed and the spacecraft radial position, an approach was adopted to separate the effects of the two variables and determine the switchback rate dependence on each. To improve the statistics of the analysis, the >3000 1 hour average steps at radial distances to 100 solar radii for orbits 3, 4, 5, 6, and 7 were aggregated into the data set after current sheet intervals were removed. The magnetic field data in these analyses were produced by the Fields instrument suite on Parker Solar Probe (Bale et al. 2016), while the wind speed came from the Sweap instrument suite (Kasper et al. 2016).

A series of Poisson regression models was developed to predict the switchback rate as functions of wind speed and solar distance (see the Appendix for details). These models assume that the logarithm of the observed switchback rate, $s$, is a
Poisson random variable with

$$\ln \mathbb{E}[s|v, d] = f(v, d),$$

where $v$ and $d$ are the solar wind speed and distance, respectively, and the expectation $\mathbb{E}[.]$ is the rate parameter of the Poisson distribution. The logarithm of the expectation was used to prevent negative values of the switchback rate from occurring. For the function $f$, polynomial functions of order $k$ with $k$ ranging from 0 to 4 were tested, where the $k$th-order function includes all terms up to order $k$, i.e.,

$$f_0(v, d) = \alpha_0$$
$$f_1(v, d) = \alpha_0 + \alpha_1 v + \alpha_2 d$$
$$f_2(v, d) = \alpha_0 + \alpha_1 v + \alpha_2 d + \alpha_3 v^2 + \alpha_4 d^2 + \alpha_5 v d.$$

Via cross validation, the order $k > 1$ models did not provide better fits than the first-order model to the data on switchback rate versus solar wind speed and radial location (see the Appendix for details). The first-order model prediction is shown in Figure 4, with the color contours indicating switchbacks/hour as a function of solar wind speed and radial distance. The faint black points represent the individual data points used to obtain the fit. This plot may be understood by imagining a horizontal line at any radial distance. Along such a line, the switchback rate increases with solar wind speed. Similarly, along a vertical line at any solar wind speed, the switchback rate is approximately constant, showing that the switchback rate does not depend on radial distance.

Figure 5. The dashed curves are Poisson linear regressions of the switchback rate on solar wind speed (panel (A)) and the radial location of the spacecraft (panel (B)). The panels show that the average switchback rate (the dashed curves) increases with solar wind speed and does not depend on the spacecraft radial position. The solid black curve in panel (A) is the contribution to the switchback rate resulting from the spacecraft sampling more radial information per hour at higher solar wind speeds. The colored points are the underlying observations, with the color indicating the radial position (panel (A)) or wind speed (panel (B)).

The Astrophysical Journal, 919:60 (10pp), 2021 September 20 Mozer et al.
Because the first-order model linearly combines solar wind speed and radial distance, one is justified in examining the marginal regression model based only on solar wind speed, \[ f(v) = a_0 + a_1 v, \] as in Figure 5(A). The dashed line indicates the resulting switchback rate as a function of solar wind speed and the colors of the data points indicate their radial locations.

**Figure 6.** The switchback rate dependence on wind speed and distance from the Sun for the five Parker Solar Probe orbits studied in this paper. In every orbit, the switchback rate increased with solar wind speed and did not depend on radial position. This shows the robustness of these results.
(The dashed line is slightly curved due to the nonlinearity of the link function in a Poisson regression.) The increase in switchback rate with solar wind speed is due in part to the fact that, at higher wind speeds, more space is surveyed by the spacecraft in each hour of switchback data. At a speed of $200\,\text{km}\,\text{s}^{-1}$ the switchback rate was about one per hour in Figure 5(A). Thus, at $600\,\text{km}\,\text{s}^{-1}$ (3 times the speed) the switchback rate due to this effect is three switchbacks/hour. This dependence is shown by the solid curve in Figure 5(A), which accounts for about half of the observed switchback increase with solar wind speed.

Similarly, the marginal regression on radial distance, with $f(r) = a_0 + a_1 r$, is presented in Figure 5(B), with the dashed curve showing that the switchback rate did not depend on distance from the Sun. Thus, Figure 5 shows that the average switchback rate increases with solar wind speed and does not depend on radial location. It is noted that there are not many points at wind speeds greater than $600\,\text{km}\,\text{s}^{-1}$ in this data set because the spacecraft encountered mostly slow solar winds during these measurements.

To show that the results of switchback dependence on wind speed and radial position are robust, the above analysis was repeated for each of the five orbits in the data set, with the results presented in Figure 6. Clearly, both the variation of switchback rate with wind speed and the lack of dependence on distance from the Sun are seen on every orbit.

A similar analysis was performed on the maximum switchback rotation angle in each hour of data. For this analysis, cross validation indicates that first- and higher-order polynomial models predict the data no more reliably than a zeroth-order (constant) model (see the Appendix for details). Figures 7 and 8 present joint and marginal first-order fits to the hourly averaged maximum rotation angle, respectively, offering little indication of a linear trend. The maximum switchback rotation angle is thus approximately independent of the solar wind speed and radial distance. The hourly averaged rotation angle is also independent of radial distance, as shown by a similar analysis.

To understand the spread of the data points in Figures 5 and 8, it is important to realize that these figures are not snapshots in time. There was up to a 28 day difference in the time between points in each of the five perihelion passes, and the five passes were separated by about 16 months. In addition, each pass covered a Carrington longitude of $\sim 90^\circ$. Thus, the data encompass a huge range of initial conditions on the Sun that produce a huge range of switchback numbers and sizes, solar wind speeds, etc. The curves of Figures 5 and 8 provide averages over all of these conditions.

There is a suggestion in Figures 4 and 5 that the number of hours without switchbacks is greatest at the lowest wind speeds. To test this possibility, Figure 9 gives the percentage of hours during which the number of switchbacks was zero or one versus the solar wind speed. At the lowest speeds, there were fewer than two switchbacks in each hour about 94% of the time. This shows that the quiet intervals between switchback patches occurred at the lowest solar wind speeds.

4. Discussion

Before comparing switchback observation rates with the two classes of source theory (transition region and in situ sources), it is necessary to consider the size variation of switchbacks with distance from the Sun because the switchback size affects the probability of observing a switchback. Because switchbacks are magnetic structures embedded in the background magnetic field, the switchback size should increase as the magnetic field decreases with distance from the Sun. As is both expected and observed in Figure 10(A), the radial magnetic field decreased as $1/R$, while the tangential field of Figure 10(B) decreased as $1/R^2$ where $R$ is the distance from the Sun. Thus, the observed switchback duration, which is proportional to its radial dimension, should increase as $R^2$ while the tangential dimension, which is proportional to the probability of detection, should increase as $R$. It is not possible to measure the tangential (perpendicular to the flow) component of the switchback size on a single spacecraft, but it is possible to test the expected $R^2$ dependence of its radial dimension by measuring the switchback durations as a function of radius. In Figure 10(C), the quantity $D/(V/400)(2500/R^2)$ is plotted versus $R$, where $D$ is the measured duration of each switchback, $V$ is the switchback speed, and $(V/400)$ normalizes the duration for the effect of different speeds. From the least-squares solid curve, it is seen that this quantity is essentially constant with distance in Figure 10(C) such that the radial dimension of switchbacks increased in size with distance in the same way that the radial magnetic field decreased. This justifies the assumption that the tangential dimension of switchbacks increased linearly with distance in the same way that the tangential field decreased. This radially decreasing tangential magnetic field decreases the density of a group of switchbacks in exactly the same way that their size increases, such that the probability of counting an existing switchback is independent of radius.

For an in situ source, the switchback rate must increase with radius because there is more space at larger distances for switchback formation. This expected increase of switchback rate for an in situ source is inconsistent with the observed constant switchback rate with distance. Thus, in situ sources are excluded by the observational data.

On the other hand, a transition region source produces switchbacks whose counting rate is independent of radial distance outside of the transition region. Thus, the constant switchback rate with radius found in this report shows that the...
Switchbacks were formed in the transition region near the Sun. The number of switchbacks observed depends on the solar wind speed, as expected because switchbacks and the solar wind are generated in the same region. Because the radial magnetic field decreases with distance more rapidly than does the tangential field, the radial dimension of switchbacks increases more rapidly with distance than the tangential dimension, resulting in switchbacks having a large and increasing aspect ratio with distance, as has been found earlier (Horbury et al. 2020; Laker et al. 2021).

The summary of this full data set is:

1. The average switchback rate did not depend on distance from the Sun.
2. The average switchback rate increased with solar wind speed.
3. The switchback tangential (to the flow) dimension increased with radial distance, $R$, while the radial dimension increased as $R^2$, resulting in a large and increasing switchback aspect ratio with distance from the Sun.
4. The hourly averaged and maximum switchback rotation angles did not depend on either the solar wind speed or the distance from the Sun.

These results are consistent with switchback formation at the transition region near the Sun. Once formed, their size and aspect ratio, but not their hourly averaged or maximum rotation angles, increased with radial distance to a distance of at least 100 solar radii. An additional result of this study is that quiet intervals between switchback patches occur at the lowest solar wind speeds.
decreased by about 20% with distance, while the extent of the switchbacks increased as the solar wind speed decreased. In panel (C), the radial component of the magnetic field decreased. In panel (B), shows the tangential component of the magnetic field increased as the radial extent of the switchbacks increased as $R^2$, the same rate that the radial component of the magnetic field decreased. In panel (C), 18 data points having durations $>40$ minutes are not included in the graph.

This work was supported by NASA contracts NNN06AA01C and NASA-G-80NSSC1. The authors acknowledge the extraordinary contributions of the Parker Solar Probe spacecraft engineering team at the Applied Physics Laboratory at Johns Hopkins University. Our sincere thanks to P. Harvey, K. Goetz, and M. Pulupa for managing the spacecraft commanding and data processing, which has become a heavy load thanks to the complexity of the instruments and the orbit. The data used in this paper are available at http://fields.ssl.berkeley.edu/data.

Appendix

We have constructed a series of Poisson regression models to predict switchbacks/hour from wind speed and solar distance. These models assume that the observed integer switchback count, $s$, is a Poisson random variable with

$$\ln \mathbb{E}[s|v, d] = f(v, d),$$

where $v$ and $d$ are the solar wind speed and distance, respectively, and the expectation $\mathbb{E}[]$ is the rate parameter of the Poisson distribution. The log transform is the standard link function used for Poisson regression and ensures the rate is nonnegative. For the function $f$, we tested models varying in complexity: polynomial functions of orders 0–4, where the $k$th-order function includes all terms up to order $k$, i.e.,

$$f_0(v, d) = \alpha_0,$$

$$f_1(v, d) = \alpha_0 + \alpha_1 v + \alpha_2 d,$$

$$f_2(v, d) = \alpha_0 + \alpha_1 v + \alpha_2 d + \alpha_3 v^2 + \alpha_4 d^2 + \alpha_5 vd,$$

etc.

The fit to the full switchback data set of models of orders 1–4 is shown in Figure A1. The coloring of the contour plot indicates the expected switchback rate. The faint black points in each graph represent the individual data points used to obtain the fit. Note the sparsity of data for high wind speeds; the fits therefore focus more on the lower wind speeds. The models of different order qualitatively agree with one another. (The fourth-order model predicts very high switchback rates in the lower right quadrant of the graph, but note that there are no data in that region.)

To evaluate each model, we split the data via three rounds of five-fold cross validation. Each split yields training and evaluation data sets with 80% and 20% of the 3000 observations, respectively. The training set is used to obtain a maximum likelihood fit of the $\alpha$ coefficients, and the evaluation set is then used to estimate two measures of model quality: the validation score, which is a constant offset from the log likelihood of the evaluation set under the model, and the Pearson correlation between the observed and predicted mean number of switchbacks.

These quality measures, shown in Figure A2 for models of orders 0–4, indicate that the first-order (linear) model provides the best and most parsimonious account of the data. This model justifies presenting the marginal regression curves in the main article.

We have also examined the relationship between $\theta_{\text{max}}$, the maximum rotation angle of the switchbacks in each hour of data, and the wind speed and solar distance, using an analogous procedure. Because $90^\circ \leq \theta_{\text{max}} \leq 180^\circ$, we used a circular regression (Sarma & Jammalamadaka 1993) that assumes that $\theta_{\text{max}}$ is a draw from a von Mises distribution with mean...
\[ \mu = l(f(v, d)) \]

where we define the inverse link function \( l \) as

\[ l(x) = \frac{1}{2} \tan^{-1}x + 135^\circ. \]

Assuming an arctan function that returns a value in \([-90^\circ, +90^\circ]\), predictions of \( \theta_{\text{max}} \) will be in the correct range with \( x \in [-\infty, +\infty] \).

In agreement with the switchback prediction, our cross-validation simulations suggest that a model of order 1 is appropriate for \( \theta_{\text{max}} \) prediction, as shown in Figure A3. As Figure A4 indicates, models of all orders have essentially the same key characteristics: little sensitivity of \( \theta_{\text{max}} \) to either wind speed or distance and no strong interactions between the two variables. Figure A3 also shows this because the validation score for the \( k = 0 \) model is no worse than any \( k \neq 0 \) model that involves wind speed and distance.
Figure A2. Validation score (related to log likelihood) and correlation are shown for models of orders 0–4. The correlation is undefined for model order 0. Error bars indicate ±1 standard error of the mean over the 15 data splits.

Figure A3. Validation score (related to log likelihood) and correlation are shown for models of orders 0–4. The correlation is undefined for model order 0. Error bars indicate ±1 standard error of the mean over the 15 data splits.
Figure A4. The expectation of $\theta_{\text{max}}$ as a function of wind speed and distance, for models of orders 1–4.

**References**

Axford, W. I., & McKenzie, J. F. 1997, in The Solar Wind, in Cosmic Winds and the Heliosphere, ed. J. R. Jokipii, C. P. Sonett, & M. S. Giampapa (Tucson, AZ: Univ. Arizona Press), 31

Bale, S. D., Badman, S. T., Bonnell, J. W., et al. 2019, *Natur*, 570, 232

Bale, S. D., Goetz, K., Harvey, P. H., et al. 2016, *SSRv*, 204, 49

Baty, H., & Keppens, R. 2002, *ApJ*, 580, 800

Borovsky, J. E. 2016, *JGRA*, 121, 5055

Chhiber, R., Usmanov, A. V., DeForest, C. E., et al. 2018, *ApJL*, 856, L39

De Pontieu, B., McIntosh, S. W., Hansteen, V., et al. 2009, *ApJL*, 701, 1

DeForest, C., Howard, R. A., Velli, M., et al. 2018, *ApJ*, 862, 18

DeForest, C. E., Matthaeus, W. H., Viall, N. M., & Cranmer, S. R. 2016, *ApJ*, 828, 66

Drake, J. F., Agapitov, O., Swisdak, M., et al. 2021, *A&A*, 650, A2

Dudok de Wit, T., Krasnoselskikh, V., Bale, S. D., et al. 2020, *ApJS*, 246, 39

Fisk, L. A. 2005, *ApJ*, 626, 563

Fisk, L. A., & Kasper, J. C. 2020, *ApJL*, 894, 1

Horbury, T. S., Woolsley, T., Laker, R., et al. 2020, *ApJS*, 246, 2

Kasper, J. C., Abiad, R., Austin, G., et al. 2016, *SSRv*, 204, 131

Kasper, J. C., Bale, S. D., Belcher, J. W., et al. 2019, *Natur*, 561, 228

Krasnoselskikh, V., Larosa, A., Agapitov, O., et al. 2020, *ApJ*, 893, 93

Laker, R., Horbury, T. S., Bale, S. D., et al. 2021, *A&A*, 650, A1

Landi, S., Hellinger, P., & Velli, M. 2006, *GeoRL*, 33, 14

Magyar, N., Utz, D., Erdelyi, R., & Nakariakov, V. 2021a, *ApJ*, 911, 75

Magyar, N., Utz, D., Erdelyi, R., & Nakariakov, V. 2021b, *ApJ*, 914, 75

Malagoli, A., Gianluigi, B., & Rossner, R. 1996, *ApJ*, 465, 708

Matteini, L., Horbury, T. S., Neugebauer, M., et al. 2014, *GeoRL*, 41, 259

McManus, M. D., Bowen, T. A., Mallet, A., et al. 2020, *ApJS*, 246, 2

Mozer, F. S., Agapitov, O. V., Bale, S. D., et al. 2020, *ApJS*, 246, 68

Ruffolo, D., Matthaeus, W. H., Chhiber, R., et al. 2020, *ApJ*, 902, 20

Samanta, T., Tian, H., Yurchyshyn, V., et al. 2019, *Sci*, 566, 890

Sarma, Y. R., & Jammalamadaka, S. R. 1993, in Statistical Science and Data Analysis: Proc. Third Pacific Area Statistical Conf., ed. K. Matsusita, M. L. Pur, & T. Hayakawa (Utrecht: VSP), 109

Squire, J., Chandran, B. D. G., & Meyer, R. 2020, *ApJL*, 891, L2

Suess, S. T. 2007, Second Solar Orbiter Workshop Report, ESA SP-66 ESA, 641

Tenerani, A., Velli, M., Matteini, L., et al. 2020, *ApJS*, 246, 2

Yamauchi, Y., Suess, S. T., Steinberg, J. T., et al. 2004, *JGRA*, 109, A3