Age-period-cohort Analysis of Healthy Lifestyle Behaviors Using the National Health and Nutrition Survey in Japan

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Objectives: This study conducted an age-period-cohort (APC) analysis of trends in healthy lifestyle behaviors in Japan.

Methods: We used National Health and Nutrition Survey data on salt intake and prevalence of smoking, drinking, and physical activity between 1995 and 2018 in Japan. Age groups were defined from 20 years to 69 years old in 10-year increments. Cohorts were defined for each age group of each year with a 1-year shift, and cohorts born in 1926-1935 (first cohort) until 1989-1998 (last cohort) were examined. We conducted a Bayesian APC analysis, calculating estimated values for each behavior by age group, period, and cohort.

Results: Estimated salt intake decreased from cohorts born in the 1930s to the 1960s, but increased thereafter in both genders, and the magnitude of increase was larger for men. Estimated smoking prevalence increased in the cohorts starting from the 1930s for men and the 1940s for women, and then decreased starting in the cohorts born in the 1970s for both genders. Although estimated drinking prevalence decreased starting in the cohorts born in approximately 1960 for men, for women it increased until the cohorts born in approximately 1970. Estimated physical activity prevalence decreased starting in the cohorts born in the 1940s in both genders, but the magnitude of decrease was larger for women.

Conclusions: Trends in cohort effects differed by gender, which might be related to changes in the social environment for women. Improvements in dietary and exercise habits are required in more recently born cohorts of both genders.

Key words: Cohort effect, Life style, Smoking, Alcohol drinking, Japan

INTRODUCTION

The increase in the number of patients with lifestyle-related diseases is a major concern in Japan, as that of major diseases such as cancer, cardiovascular diseases (CVD), diabetes, and dyslipidemia in Japan has consistently increased over the years [1]. According to the Patient Survey in Japan [1], the estimated number of cancer patients increased from 1.26 million in 1999 to 1.73 million in 2017. In addition, the estimated number of patients with CVD increased from 10.87 million in 1999 to 13.01 million in 2017. As a result, national medical expenses have continued to rise; according to national medical expenses data [2], the expenses for CVDs amounted to 6.08 trillion yen in 2017, those for cancer amounted to 3.82 trillion yen, and those for endocrine and metabolic diseases like diabetes amounted to 2.09 trillion yen. Therefore, better prevention of lifestyle-related diseases is required in Japan.

Dietary habits, smoking, drinking, and physical activity (PA) are often pointed out as lifestyle choices that trigger lifestyle-related diseases [3-5]. Among dietary habits, fat and salt intake are often considered as risk factors for CVDs or dyslipidemia [6,7]. The association between these lifestyle factors and
disease has been clearly demonstrated in many studies. It is also important, though, to investigate changes in lifestyle behaviors. Salt intake and the prevalence of smoking have decreased over the past 10 years among both men and women [8]. Conversely, the prevalence of alcohol drinking (hereafter referred to as drinking prevalence) has remained stable among men and has been increasing among women [8]. Although step counts have generally remained stable for both men and women [8], a previous report indicated that average step counts decreased between before 2000 and 2010 for both men and women [9]. Although studies have analyzed changes in lifestyle behaviors, analyses focusing on cohort effects on lifestyle behaviors are still rare in Japan, and changes in lifestyle behaviors among cohorts are uncertain.

Age-period-cohort (APC) analysis is often used as a method of decomposing statistics into age, period, and cohort effects [10]. APC analysis has been used to analyze trends in cancer and cardiovascular mortality in Japan [11,12], and reasons for changes in cohort effects on diseases have been discussed, taking into account lifestyle changes among cohorts. In order to interpret the results of these studies estimating cohort effects on lifestyle-related diseases, it is also meaningful to conduct APC analyses for disease risk factors. Changes among cohorts in the energy intake ratio from each type of nutrient have previously been investigated using APC analysis in Japan [13]. However, APC analyses of cohort effects on salt intake and/or the prevalence of smoking, drinking, and PA have not yet been conducted. Therefore, in this study, we conducted an APC analysis of lifestyle behavior trends.

**METHODS**

We used data from the National Health and Nutrition Survey in Japan [14,15]. This survey is conducted every year to obtain detailed information on the physical status, nutritional intake, and lifestyles of Japanese people [8]. Subjects are chosen by stratified random sampling from all households in Japan. Subjects are asked by questionnaire about their physical conditions, including height and weight, nutrient intake, and lifestyle information including PA and sleep. The questionnaire about nutrient intake measures levels of consumption of various nutritional sources, including total energy, protein, fat, cholesterol, carbohydrate, dietary fiber, vitamin, salt, potassium, and calcium [14,15]. We analyzed data on salt intake, smoking prevalence, drinking prevalence, and the prevalence of PA (hereafter referred to as PA prevalence) by age group and gender between 1995 and 2018. Smoking prevalence and drinking prevalence were defined, respectively, the rate of persons who habitually smoked or drank at a given time. PA prevalence was defined as the proportion of subjects who exercised for more than 30 minutes at a time, at least twice per week, and who continued this habit for more than 1 year. Salt intake was defined as the mean intake of salt per day. Since salt intake is known to be affected by total energy intake, we analyzed mean salt intake (g) per 1000 kcal [16]. Drinking prevalence was not measured in 2013. Age groups were defined using 10-year increments from 20-29 years to 60-69 years. Those who were 60-69 years old in 1995 (i.e., those born between 1926 and 1935) were the first cohort, as they were the oldest birth cohort in the data set. Through a 1-year shift of birth years, starting from the first cohort, the age group of participants who were 20-29 years old in 2018 (i.e., those born between 1989 and 1998) comprised the last cohort.

First, we calculated birth cohort-specific salt intake, smoking prevalence, drinking prevalence, and PA prevalence by age group for men and women to describe cohort effects. We also calculated age-adjusted rates or amounts of the lifestyle behaviors. Age-adjusted rates or amounts of lifestyle behaviors were calculated using the direct method [17], and the total population in 1995 was used as the standard population for both men and women. Data from Vital Statistics in Japan [18] were used for population ratios. We then used a Bayesian APC model. To analyze salt intake, we used a Bayesian APC model that assumed that salt intake values followed a normal distribution whose mean was an additive model of APC effects, similar to the model used by Kwok et al. [19]. Let $y_{ij}$ be the salt intake for the age group $i$ ($1, \ldots, I$) in year $j$ ($1, \ldots, J$). In the model, salt intake is assumed to follow the normal distribution shown below, whose mean is $\lambda_{ij}$:

$$y_{ij} \sim \text{Normal}(\lambda_{ij}, \sigma^2),$$

$$\lambda_{ij} = \delta + \alpha_i + \beta_j + \gamma_k$$

Where $\delta$ is the intercept, $\alpha_i$ is the age group effect, $\beta_j$ is the period effect, $\gamma_{k}$ ($k = 1, \ldots, K$) is the cohort effect, and $\sigma^2$ is the variance of $y_{ij}$. $I$, $J$, and $K$ are the number of time points for age groups, periods, and cohorts, respectively. To identify each parameter, the sum of each effect was restricted to zero, and a first-order random-walk was used as a prior for each effect [20]. Because smoking prevalence, drinking prevalence, and...
Figure 1. Birth cohort-specific salt intake, smoking prevalence, drinking prevalence, and physical activity (PA) prevalence by age groups for both men and women.
PA prevalence were all proportion data, to analyze these we used a Bayesian APC model that assumed that the log odds of each prevalence, \( \eta_{ij} = \log(y_{ij}/(1-y_{ij})) \), followed a normal distribution whose mean was an additive model of APC effects. The estimated prevalence or amount of each lifestyle behavior was then calculated for each age group, period, and cohort using the estimated parameters. For example, the estimated prevalence for age group \( i \) can be calculated as \( \expit(\hat{\delta} + \hat{\alpha}_i) \), where \( \hat{\delta} \) and \( \hat{\alpha}_i \) are the estimated values for the intercept \( \delta \) and the age effect \( \alpha_i \), and \( \expit \) is the sigmoid function.

We investigated the increase or decrease of estimated effects based on the effect estimates, as is usually done in APC analyses [11-13,21,22]. All statistical analyses were conducted using R version 3.5.1 (https://cran.r-project.org/bin/windows/base/old/3.5.1/), and the Hamiltonian Monte Carlo method was used to estimate the parameters of the APC model [23].

**Table 1.** Yearly age-adjusted salt intake, smoking prevalence, drinking prevalence, and physical activity (PA) prevalence from 1995 to 2018

| Year | Salt intake\(^1\) | Smoking prevalence | Drinking prevalence | PA prevalence | Salt intake\(^1\) | Smoking prevalence | Drinking prevalence | PA prevalence |
|------|-------------------|--------------------|-------------------|---------------|-------------------|--------------------|-------------------|---------------|
| 1995 | 6.4               | 56.7               | 54.9              | 24.7          | 6.9               | 11.8               | 7.8               | 20.7          |
| 1996 | 6.5               | 54.1               | 54.0              | 24.0          | 6.9               | 10.9               | 8.7               | 21.6          |
| 1997 | 6.3               | 57.5               | 51.6              | 26.5          | 6.8               | 13.7               | 10.0              | 23.5          |
| 1998 | 6.2               | 56.0               | 52.0              | 22.9          | 6.8               | 12.9               | 10.6              | 22.4          |
| 1999 | 6.2               | 53.9               | 52.4              | 27.7          | 6.7               | 12.5               | 9.0               | 25.2          |
| 2000 | 6.1               | 53.5               | 50.5              | 28.1          | 6.7               | 14.4               | 10.3              | 24.5          |
| 2001 | 5.8               | 53.0               | 52.9              | 27.0          | 6.3               | 12.2               | 10.8              | 23.7          |
| 2002 | 5.8               | 50.1               | 47.0              | 26.6          | 6.3               | 13.4               | 10.2              | 24.1          |
| 2003 | 5.7               | 52.3               | 39.0              | 24.5          | 6.1               | 14.3               | 8.5               | 20.9          |
| 2004 | 5.5               | 48.8               | 39.1              | 23.0          | 5.9               | 14.4               | 8.6               | 22.2          |
| 2005 | 5.6               | 45.1               | 38.4              | 23.1          | 6.0               | 14.9               | 9.5               | 23.3          |
| 2006 | 5.6               | 45.5               | 36.9              | 24.4          | 5.9               | 13.0               | 8.6               | 23.8          |
| 2007 | 5.4               | 45.9               | 37.4              | 23.6          | 5.8               | 14.1               | 10.1              | 21.3          |
| 2008 | 5.5               | 43.6               | 36.6              | 26.3          | 5.8               | 12.3               | 8.1               | 21.8          |
| 2009 | 5.4               | 44.0               | 36.3              | 26.6          | 5.6               | 13.9               | 8.8               | 21.5          |
| 2010 | 5.3               | 37.7               | 34.7              | 27.6          | 5.7               | 11.4               | 8.7               | 21.4          |
| 2011 | 5.3               | 38.3               | 36.5              | 26.1          | 5.7               | 12.8               | 10.2              | 22.6          |
| 2012 | 5.2               | 39.7               | 33.8              | 26.0          | 5.5               | 11.5               | 8.8               | 19.5          |
| 2013 | 5.1               | 39.0               | -                 | 22.1          | 5.4               | 11.3               | -                 | 19.8          |
| 2014 | 4.9               | 39.1               | 33.3              | 21.5          | 5.4               | 11.7               | 10.3              | 16.3          |
| 2015 | 5.0               | 35.5               | 32.3              | 24.1          | 5.2               | 9.8                | 9.7               | 18.7          |
| 2016 | 5.0               | 36.6               | 32.6              | 24.9          | 5.3               | 10.7               | 10.4              | 18.2          |
| 2017 | 4.9               | 34.1               | 32.2              | 27.0          | 5.1               | 9.0                | 10.1              | 18.6          |
| 2018 | 5.0               | 33.2               | 31.6              | 21.7          | 5.3               | 10.6               | 10.4              | 16.0          |

Values are presented as percentage.
\(^1\)Unit: g/1000 kcal.

**Ethics Statement**

This study did not require ethical approval because it used public data.

**RESULTS**

Figure 1 shows results for birth cohort-specific salt intake, smoking prevalence, drinking prevalence, and PA prevalence by age group for men and women. Salt intake decreased across cohorts in all age groups for men and women. Although smoking prevalence decreased in all age groups across cohorts for men, smoking prevalence did not decrease in women who were 40-69 years of age. Although drinking prevalence in all age groups decreased for men, the trends differed across age groups for women. Although PA prevalence remained relatively stable across 20-59 years of age in men, it showed de-
Figure 2. Results of age-period-cohort analysis of salt intake (unit: g/1000 kcal), smoking prevalence, drinking prevalence, and physical activity (PA) prevalence for men. Solid lines signify estimates of each effect, and the shadings show 95% credible intervals of each effect.
Figure 3. Results of age-period-cohort analysis of salt intake (unit: g/1000 kcal), smoking prevalence, drinking prevalence, and physical activity (PA) prevalence for women. Solid lines signify estimates of each effect, and the shadings show 95% credible intervals of each effect.
increasing trends in some age groups for women.

Table 1 shows yearly age-adjusted salt intake, smoking prevalence, drinking prevalence, and PA prevalence from 1995 to 2018. Age-adjusted salt intake decreased throughout the study period for both men and women. Age-adjusted smoking prevalence decreased throughout the study period for men, and for women, it showed a decreasing trend starting in approximately 2005. Although age-adjusted drinking prevalence decreased throughout the study period for men, for women it remained relatively stable. Age-adjusted PA prevalence continued to fluctuate for both men and women, and it decreased in recent years for women.

Figure 2 shows results of the APC analysis of salt intake, smoking prevalence, drinking prevalence, and PA prevalence for men. Estimated salt intake increased until 50-59 years old, and it decreased throughout the study period. In addition, the estimated salt intake decreased by approximately 0.6 g in the cohorts born between approximately 1930 and the mid-1960s, and it increased by approximately 0.3 g in the cohorts born in roughly 1990. Estimated smoking prevalence decreased both throughout the study period and with increasing age. In addition, estimated smoking prevalence increased from approximately 0.37 to 0.48 in the cohorts born between the mid-1930s and the late 1970s and began to decrease thereafter. The peak estimated drinking prevalence was at 40-59 years old. Estimated drinking prevalence decreased in the early 2000s, and showed a slight decreasing trend thereafter. In addition, estimated drinking prevalence began to decrease starting in the cohorts born in approximately 1960, and it decreased from approximately 0.45 to 0.23 in the cohorts born between approximately 1960 and 1990. Estimated PA prevalence began to increase starting in the age groups of 30-39 years, and showed a stable trend over the years. In addition, estimated PA prevalence increased until the cohorts born in approximately 1940, and showed slight decreasing trend thereafter. It decreased from 0.28 to 0.21 in the cohorts born between approximately 1940 and 1990.

Figure 3 shows the results of the APC analysis of salt intake, smoking prevalence, drinking prevalence, and PA prevalence for women. Peak estimated salt intake was found at 50-69 years old, and it decreased throughout the study period, similar to results observed in men. In addition, estimated salt intake decreased by approximately 0.9 g in the cohorts born between the 1930s and the mid-1970s, and slightly increased thereafter. Estimated smoking prevalence decreased in older age groups, and it started to decrease in 2005. In addition, the estimated smoking prevalence increased from approximately 0.08 to 0.15 in the cohorts born between the 1940s and the mid-1970s and decreased to approximately 0.10 in the cohorts born in approximately 1990. The peak estimated drinking prevalence was found at 40-49 years old. Estimated drinking prevalence also remained relatively stable over the years, although it slightly increased in recent years. In addition, estimated drinking prevalence increased from 0.06 to 0.11 in the cohorts born between approximately 1940 and the mid-1970s, and decreased thereafter. Estimated PA prevalence tended to increase with increasing age, and slightly decreased starting in approximately 2000. In addition, estimated PA prevalence increased until cohorts born in the mid-1940s, and decreased thereafter. It decreased from approximately 0.25 to 0.14 in the cohorts born between the 1940s and approximately 1990.

**DISCUSSION**

The trends in cohort effects on salt intake were similar in both genders. A decrease in salt intake was observed in both genders in cohorts born in the 1930s-1960s. It has been shown that systolic blood pressure continued to decrease between 1960 and 1990 in Japan [24], and decreased salt intake across periods and cohorts is thought to be related to this result. However, our results suggest that salt intake is increasing in more recently born cohorts for both men and women, particularly for men.

Regarding cohort effects on smoking prevalence, estimated smoking prevalence increased in the cohorts born from approximately the 1930s to the 1970s and began to decrease thereafter for men. It is known that smoking prevalence among Japanese people began to decrease from approximately the late 1960s onwards [25], and it is suggested that decrease of the cohort effect occurred later than the overall decrease in smoking prevalence. The cohort effect on heart disease mortality showed a decrease starting in cohorts born in approximately 1970 for men [26], which might be related to the cohort effect of decreasing smoking prevalence. In contrast, estimated smoking prevalence for women increased until the cohorts born in the mid-1970s. The rate of the decrease of smoking prevalence for women was smaller than that for men [25], which is considered to be related to the obvious increase of smoking prevalence over the cohorts among wom-
en. According to an epidemiological study investigating smoking status and socioeconomic factors in Japan [27], being employed was strongly associated with smoking in women than in men. Therefore, the increase in the employment rate of women across cohorts might be related to their increased smoking prevalence.

Regarding cohort effects on drinking prevalence, a difference was observed in the turning point between men and women. Although estimated drinking prevalence for men began to decrease starting in the cohorts born in approximately 1960, for women the decrease started in the cohorts born in approximately 1970. The trend in the cohort effect on drinking prevalence was similar to that of smoking prevalence in women. As shown in Table 1, age-adjusted drinking prevalence increased among women, unlike men, in recent years, and the difference in these trends is considered to contribute to the results of age-adjusted drinking prevalence. Although the underlying reason for this difference is uncertain, it might be related to changes in the social environment of women among cohorts, as previous studies [28,29] have shown associations between employment status and drinking in other countries. Japanese people often consume alcohol in order to form friendly relationships with others [30], which is thought to be related to employment status. Therefore, a study investigating predictors of drinking behavior in Japan may be needed.

Although estimated PA prevalence slightly decreased in men starting in the cohorts born in the 1940s, it was found to consistently decrease in women starting in the cohorts born in approximately 1940. A possible reason for the obvious decrease of the cohort effect for women is the employment rate. The employment rate of women has continued to increase over the cohorts, and sedentary time has therefore increased as well. It has been reported in other countries that women with full-time sedentary jobs experienced less light-intensity and lifestyle-intensity activity than healthy non-workers [31], and also that sedentary behavior was particularly common among white-collar women [32]. A similar phenomenon may therefore be observed in Japan. According to a previous APC analysis, an increase in cohort effect for obesity prevalence in women in Japan began to emerge starting approximately in the 1960s [13], and the decrease in PA prevalence across the cohorts is thought to be related to this result. Furthermore, the cohort effect on mortality rate of cerebrovascular diseases and ischemic heart diseases was shown to have increased continuously in women starting in the cohorts born in approximately 1960 [26], which might also be related to the cohort effects on decreased PA prevalence and the cohort effects on salt intake.

To conclude, in this study, we estimated cohort effects for each lifestyle behavior. The results from this study will be useful when interpreting results of APC analyses of lifestyle-related diseases in the future. Men showed higher estimated smoking and drinking prevalence in every age, period, and cohort, and a preventive approach to smoking and drinking is still required, particularly for men. On the other hand, for women, estimated PA prevalence clearly decreased over the cohorts, meaning that greater public awareness of insufficient PA is required in more recently born cohorts. Finally, estimated salt intake increased in recently born cohorts for both men and women, particularly for men. In light of the finding that the proportion of energy intake from fat has continuously increased for both men and women [13], changes in dietary habits are required for both genders.

**CONFLICT OF INTEREST**

The author has no conflicts of interest associated with the material presented in this paper.

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**AUTHOR CONTRIBUTIONS**

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