Sterrett Procedure for the Generalized Group Testing Problem

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July 3, 2018

Abstract

Group testing is a useful method that has broad applications in medicine, engineering, and even in airport security control. Consider a finite population of \( N \) items, where item \( i \) has a probability \( p_i \) to be defective. The goal is to identify all items by means of group testing. This is the generalized group testing problem. The optimum procedure, with respect to the expected total number of tests, is unknown even in case when all \( p_i \) are equal. Hwang (1975) proved that an ordered partition (with respect to \( p_i \)) is the optimal for the Dorfman procedure (procedure \( D \)), and obtained an optimum solution (i.e., found an optimal partition) by dynamic programming. In this paper, we investigate the Sterrett procedure (procedure \( S \)). We provide close form expression for the expected total number of tests, which allows us to find the optimum arrangement of the items in the particular group. We also show that an ordered partition is not optimal for the procedure \( S \) or even for a slightly modified Dorfman procedure (procedure \( D' \)). This discovery implies that finding an optimal procedure \( S \) appears to be a hard computational problem. However, by using an optimal ordered partition for all procedures, we show that procedure \( D' \) is uniformly better than procedure \( D \), and based on numerical comparisons, procedure \( S \) is uniformly and significantly better than procedures \( D \) and \( D' \).

Keywords: Cost function; Information theory; Partition problem
1 Introduction

In 1943, Robert Dorfman introduced the concept of group testing as a need to administer syphilis tests to millions of individuals drafted into the U.S. army during World War II. The test for syphilis was a blood test called the Wassermann test (Wassermann et al. 1906). The nice description of the Dorfman (1943) procedure is given by Feller (1950): "A large number, \( N \), of people are subject to a blood test. This can be administered in two ways. (i) Each person is tested separately. In this case \( N \) tests are required. (ii) The blood samples of \( k \) people can be pooled and analyzed together. If the test is negative, this one test suffices for the \( k \) people. If the test is positive, each of the \( k \) persons must be tested separately, and all \( k + 1 \) tests are required for the \( k \) people. Assume the probability \( p \) that the test is positive is the same for all and that people are stochastically independent."

Procedure (ii) is commonly referred to as the Dorfman two-stage group testing procedure.

Since the Dorfman work, the group testing has widespread applications. The partial list of applications includes blood screening to detect human immunodeficiency virus (HIV), the detection of hepatitis B virus (HBV), and other diseases (Gastwirth and Johnson 1994; Bar-lev et al. 2010; Bilder et al. 2010; Stramer 2011), for screening chemical compounds as part of the drug discovery (Zhu et al. 2001), DNA screening (Macula 1999; Du and Hwang 2006; Cao and Sun 2016), quality control in product testing (Sobel and Groll 1959; Bar-lev et al. 1990), and communication networks (Wolf 1985).

There is a logical inconsistency in the Dorfman two-stage group testing procedure (procedure \( D \)). It is clear that any "reasonable" group testing plan should satisfy the following property: "A test is not performed if its outcome can be inferred from previous test results" (Ungar 1960, p. 50). The procedure \( D \) does not satisfy this property, since if the group is positive and all but the last person are negative, the last person is still tested. The modified Dorfman procedure (Sobel and Groll 1959) (defined as \( D' \)) would not test the last individual in this case. Even though this improvement is small with respect to the expected number of tests, it has a large impact on optimal design (Malinovsky and Albert 2016).

Sterrett (1957) suggested an improvement of the procedure \( D' \) in the following way. If in the first stage of the procedure \( D' \) the group is infected, then in the second stage
individuals are tested one-by-one until the first infected individual is identified. Then the first stage of procedure is applied to the remaining (non-identified) individuals. The procedure is repeated until all individuals are identified. The intuition behind Sterrett procedure (procedure $S$) is the following: if $p$ is small, then the probability that two or more subgroups are positive is very small and it will bring the advantage of procedure $S$ over $D'$. On the other hand, the procedures $D, D'$ are very simple and the second stage can be performed (if there is a need) simultaneously, but procedure $S$ is sequential one-by-one and, therefore, more time consuming. Therefore, practical application depends on a real need. The recent results, comparisons, and review of these procedures can be found in Malinovsky and Albert (2016).

There are only a few fundamental results in binomial group testing that provide insights on the structure of $E(N, p)$, the minimum expected number of tests, which generally is not known. The most important one is due to Ungar (1960). Ungar characterized the optimality of any group testing algorithm and proved that if $p \geq p_U = \frac{3 - 5^{1/2}}{2} \approx 0.38$, then there does not exist an algorithm that is better than individual one-by-one testing, i.e., $E(N, p) = N$, for $p \geq p_U$.

To the best of our knowledge, the generalized group testing problem was introduced, in the first time, by Nebenzahl and Sobel (1973). In the motivation problem, the units arrived in the groups of different sizes: group $i$ has size $i$. Therefore, the probability that the group $i$ is defective, (i.e., include at least one defective unit) is $1 - q^i$, $q = 1 - p$. Then group $i$ can be treated as a single unit with $p_i = 1 - q^i$.

This work deals with the generalized group testing problem (GGTP): $N$ stochastically independent units, where unit $i$ has the probability $p_i$ ($0 < p_i < 1$) to be defective. All units has to be classified as good or defective by group testing. The group testing procedure $A$ is better (not worse) than procedure $B$, if total expected number of tests under the procedure $A$ is less (less or equals) than under the procedure $B$ for any set $\{p_1, p_2, \ldots, p_N\}$. It is important to note that even under equal probabilities case an optimum procedure is unknown (for the discussion see Malinovsky and Albert (2016)). We assume that the probabilities $\{p_1, p_2, \ldots, p_N\}$ are known and we can decide about the order in which the people will be tested.
For the GGTP, we redefined the group testing procedures $D$, $D'$, and $S$ to be partition of the finite population of size $N$ into any number of disjoint groups/subsets. Then procedure $A$ ($A \in \{D, D', S\}$) is performed on each of these groups.

Ideally, under procedure $A$ ($A \in \{D, D', S\}$) we are interested in finding an optimal partition $\{m_1, \ldots, m_I\}$ with $m_1 + \ldots + m_I = N$ for some $I \in \{1, \ldots, N\}$ such that the total expected number of tests is minimal, i.e.,

$$\{m_1, \ldots, m_I\} = \arg \min_{n_1, \ldots, n_J} E_A(n_1, n_2, \ldots, n_J),$$

subject to $\sum_{i=1}^J n_i = N, \ J \in \{1, \ldots, N\}$, \hspace{1cm} (1)

where $E_A (n_1, n_2, \ldots, n_J) = E_A (1 : n_1) + E_A (1 : n_2) + \ldots + E_A (1 : n_J)$, and $E_A (1 : n_j)$ is the total expected number of tests (under procedure $A$) in the group of size $n_j$.

It is important to note that from Ungar’s fundamental result (discussed above), it follows that if $p_i \geq p_U = (3 - 5^{1/2})/2 \approx 0.38$ for any $i = 1, \ldots, N$ then does not exist a group testing algorithm which is better than a one-by-one individual testing. In this case, the solution of (1) is $n_i = 1, i = 1, \ldots, I, I = N$.

Potentially, the optimization problem (1) can be solved under brutal search, i.e., evaluating the right-hand side of the first equation in (1) for any possible partition and then choosing an optimal one. But this task is a hard computational problem and it is impossible to perform because the total number of possible partitions of a set of size $N$ is the Bell number $B(N) = \left\lceil \frac{1}{e} \sum_{j=1}^{2N} \frac{j^N}{j!} \right\rceil$ (Bell (1934), Bollobás (2006)), which grows exponential with $N$. For example, $B(3) = 5, B(5) = 52, B(10) = 115,975, B(13) = 27,644,437$.

However, for some cost functions (in our case the cost function is $E_A (n_1, n_2, \ldots, n_J)$) a solution can be obtained with polynomial in $N$ computation cost (see, e.g., Hwang (1975); Tanaev (1979); Hwang (1981); Chakravartty et al. (1982)).

In the group-testing setting, Hwang (1975, 1981) proved that under procedure $D$ an optimal partition is an ordered partition (i.e., each pair of subsets has the property that the numbers in one subset are all great or equal to every number in the other subset). The total number of ordered partition of a set of size $N$ is $2^{N-1}$. The brutal search among all $2^{N-1}$ possibilities is still a computational intractable problem (Garey and Johnson, 1979).
But in this case, Hwang (1975) showed the existence of a polynomial time algorithm with the computational effort $O(N^2)$. This algorithm is the dynamic programming algorithm (Bellman, 1957).

Bilder et al. (2010) applied procedure $S$ to chlamydia and gonorrhea testing in Nebraska. For each covariate (gender and specimen) combination, they randomly assigned the individuals to groups of constant size. Then they estimated the individual risk probabilities $p_i$ in each group using a logistic model. The procedure $S$ was applied to each group. The individuals’ insight of the particular group were tested (if necessary) according to the decreasing order of $p_i$, i.e., first, the individual with the largest $p_i$ value, next the second largest, and so on. They assumed that the diagnostic tests are not error-free. Generally, this assumption is realistic and has to be taken into account. We do not attempt to investigate erroneous tests in the current work and discuss this as an important direction for the future investigation.

To the best of our knowledge, no optimality result is available to the other group testing procedures for the generalized problem, including procedure $S$. It can be explained by the fact that the close form expression for the $E_S(1:k)$ was not available and it is essential.

In this work, we find the closed-form expression and also show how items should be arranged in the particular group such that $E_S(1:k)$ will be minimal. We also show that for the both procedures $D'$ and $S$ an optimal partition is not ordered. Therefore, there is no available optimum solution here with polynomial in $N$ computation effort, and it seems to be a computationally hard problem. But, despite this negative answer, the performance of the procedure $S$ under an optimal ordered partition (which is not optimal, but can be obtained with $O(N^2)$ computational effort) is significantly better than the performance of an optimal procedure $D$. The comparisons with procedure $D'$ will be made also.

## 2 Basic Characteristics of the Procedures

### 2.1 Procedure $S$

In order to investigate the Sterrett procedure, we have to find the expected number of tests, in the group of size $k$, under this procedure. First, we present below two examples for the
particular cases \( k = 2 \) and \( k = 3 \). It will help to develop the intuition for the general case.

**Example 1** (Group of size \( k = 2 \)).

The tree above represents Sterrett procedure for the group of size \( k = 2 \). Starting from the root of the tree with two individuals, we perform the procedure and stop at terminal node when all (here 2) individuals are identified. The left branch of the tree represents the negative test result, and the right branch represents the positive test result. The total number of tests \( T \) with the corresponding probabilities are presented in the terminal nodes. From the tree, we obtain the expected total number of tests.

\[
E_S(1:2) = q_1q_2 + 2q_1(1-q_2) + 3(1-q_1) = 3 - q_1 - q_1q_2.
\] (2)

From Equation (2), it follows that arranging the probabilities in decreasing order \( q_1 \geq q_2 \) we obtain the minimum value of \( E_S(1:2) \).

The next example for the case \( k = 3 \) will help us to approach the general case.

**Example 2** (Group of size \( k = 3 \)).
The expected total number of tests is

\[ E_S(1:3) = 1q_1q_2q_3 + 3q_1q_2(1-q_3) + 4q_1(1-q_2) + (2 + E(2:3))(1-q_1) \]

\[ = 5 - q_1 - q_2 - q_2q_3 - q_1q_2q_3. \]  

(3)

From Equation (3), it follows that arranging the probabilities in the order \( q_2 \geq q_1 \geq q_3 \) we obtain the minimum value of \( E_S(1:3) \).

These two examples are induction steps for the general case which is presented below.

**Result 1.** (i) The total expected number of test under generalized Sterrett procedure in the group of size \( k \) (\( k \geq 1 \)) is

\[ E_S(1:k) = (2k-1) - [(q_1 + q_2 + \ldots + q_{k-1}) + q_{k-1}q_k + q_{k-2}q_{k-1}q_k + \ldots + q_1q_2 \ldots q_k]. \] \hspace{1cm} (4)

(ii) For \( k \geq 2 \), the minimum of \( E_S(1:k) \) is obtained under the order

\[ q_{k-1} \geq q_{k-2} \geq \ldots \geq q_1 \geq q_k. \] \hspace{1cm} (5)

For the proof of Result 1, see Appendix A.

**Corollary 1.** If \( p_1 = \ldots = p_N = p \), then

\[ E_S(1:k) = 2k - (k-2)q - \frac{1-q^{k+1}}{1-q}, \quad k \geq 1. \] \hspace{1cm} (6)

It is an interesting to note that in the original work of Sterrett (1957) the closed form (6) was not provided. For the discussion, please see Malinovsky and Albert (2016).

### 2.2 Procedures \( D \) and \( D' \)

In this section, we will compare the procedure \( S \) with the procedures \( D \) and \( D' \). For this purpose, we need to present some characteristics for the procedures \( D \) and \( D' \). Below we present the expected total number of tests in the group of size \( k \) for the procedures \( D \) and \( D' \).

**Procedure \( D \)**
For $k \geq 2$, the total number of tests is 1 with probability $\prod_{i=1}^{k} q_i$ and $k + 1$ with probability $1 - \prod_{i=1}^{k} q_i$. Therefore,

$$E_D(1 : k) = 1 + \left( k - k \prod_{i=1}^{k} q_i \right) 1_{\{k \geq 2\}}.$$  

(7)

**Procedure $D'$**

For $k \geq 2$, the total number of tests is 1 with probability $\prod_{i=1}^{k-1} q_i$, $k$ with probability $\prod_{i=1}^{k} q_i (1 - q_k)$, and $k + 1$ with probability $1 - \prod_{i=1}^{k} q_i - \prod_{i=1}^{k-1} q_i (1 - q_k)$. Therefore,

$$E_{D'}(1 : k) = 1 + \left( k - k \prod_{i=1}^{k} q_i - k - \prod_{i=1}^{k-1} q_i (1 - q_k) \right) 1_{\{k \geq 2\}}.$$  

(8)

From Equations (7) and (8), we obtain three important observations:

(i) In the procedure $D$, the order in which items are tested in the particular group is not important, the value of $E_D(1 : k)$ is the same for any permutation. It is not so under procedure $D'$. In order to minimize $E_{D'}(1 : k)$, the last tested item, say item $k$, should correspond to the smallest $q_i$ in the group.

(ii) From Equations (7) and (8), it obviously follows that

$$E_{D'}(1 : k) \leq E_D(1 : k).$$

(iii) From the definitions of the procedures $D'$ and $S$, it follows that they are equivalent for $k = 2$, i.e., $E_{D'}(1 : 2) = E_S(1 : 2)$. It also follows from Equations (2) and (8).

3 Finding an Optimal Procedures $D'$ and $S$ Appears to be a Hard Computational Problem

It was mentioned in the Introduction that for the arbitrary cost function, the solution of (1) is computationally hard problem (see, e.g., Garev and Johnson (1979)).
We also mentioned that Hwang (1975) had proved that an optimal ordered partition is an optimal for the procedure $D$. In the conclusions to his work, Hwang (1975) wrote: “A Dorfman procedure has the property that a unit is classified as a defective only after an individual test finds it so. There are procedures which allow a unit to be classified as defective by deduction, e.g., if a group of size $g$ is found defective and subsequent tests find $g - 1$ members of this group good, then the remaining member is classified as a defective without further testing. The savings of cost due to this possible deduction is small, but the mathematics to obtain optimal procedures would be extremely complicated. Therefore, we do not pursue this line.” It is clear that procedures $D'$ and $S$ satisfies the “deduction” property mentioned by Hwang. Even though this improvement is small with respect to the expected number of tests, it has a large impact on optimal design. It was explained in Malinovsky and Albert (2016) for the case $p_1 = \ldots = p_N$.

The following example shows that the optimal ordered partition is not optimal for the procedures $D'$ and $S$.

**Example 3.** Take $N = 4$ and $\{q_1, q_2, q_3, q_4\} = \{0.6, 0.6, 0.99, 0.99\}$. There are 8 possible ordered partitions. The optimum one among them can be found using (4) and (8) through direct calculation or using dynamic programming algorithm (Appendix B). Such optimum ordered partition is the same for both $S$ and $D'$: $\{0.6\} \cup \{0.6, 0.99, 0.99\}$ with the corresponding $E_S = 2.83794$ and $E_{D'} = 2.8438$.

Now, consider the following unordered partition $\{0.6, 0.99\} \cup \{0.6, 0.99\}$. For this partition, we have $E_S = E_{D'} = 2.832$ showing non-optimality of the optimal ordered partition for both procedures.

The above example shows that finding an optimal solution for (1) seems to be computationally hard problem under procedures $D'$ and $S$.

Another interesting observation is the following. As we already mentioned $E_{D'}(1:2) = E_S(1:2) = 3 - q_1 - q_1 q_2$. For $N = 4$ without loss of generality, let us assume $q_1 \geq q_2 \geq q_3 \geq q_4$. Suppose that we perform procedure $S$ (or $D'$) to the following ordered partition $\{q_1, q_2\} \cup \{q_3, q_4\}$. By interchanging $q_2$ with $q_3$ (and creating unordered partition), we always reduce the expected total number of tests.
3.1 Comparisons Among the Procedures

3.1.1 Comparisons

The following result is useful for the comparison.

Result 2. Denote \( \{m_1, \ldots, m_I\} \) an optimal partition under procedure \( D \), and \( \{m'_1, \ldots, m'_I\} \) an optimal ordered partition under procedure \( D' \). From Equations (7), (8), and Hwang (1975, 1981) (which proved that an optimal partition under procedure \( D \) is ordered), the following relations immediately follows:

\[
E_{D'} \left( m'_1, \ldots, m'_I \right) \leq E_{D'} \left( m_1, \ldots, m_I \right) \leq E_D \left( m_1, \ldots, m_I \right).
\] (9)

Denote \( \{m_1^*, \ldots, m_I^*\} \) an optimal ordered partition under Sterrett procedure. We are going to compare \( E_S \left( m_1^*, \ldots, m_I^* \right) \) with \( E_{D'} \left( m'_1, \ldots, m'_I \right) \).

The comparisons will be done in the following manner. We generate the vector \( p_1, p_2, \ldots, p_{100} \) from Beta distribution with parameters \( \alpha = 1, \beta > 0 \) such that \( \frac{1-p}{p} = \beta \), i.e., expectation equals to \( p \), and standard deviation equals to \( p \sqrt{\frac{1-p}{1+p}} \) (second column of Table 1). We repeat this process \( M = 1000 \) times for each value of \( p \). Each time a solution of the optimization problem (1) under ordered partition restriction for procedures \( D, D', \) and \( S \) was obtained using dynamic programming (DP) algorithm, which is presented in Appendix B. From DP algorithm we also obtained the expected total number of tests for an optimal ordered partition in procedures \( D, D', \) and \( S \).

In the following table, we present a mean (over 1000 repetitions) of the expected total number of tests under these three procedures and in the bracket we present the standard deviation of the mean.
| p   | std   | \(N = 100\) |   |   |   |
|-----|-------|-------------|---|---|---|
| 0.001 | 0.001 | 5.7519 (0.009) | 5.7383(0.009) | 3.7453(0.006) | 1.0807(0.003) |
| 0.01  | 0.0099| 17.47(0.0279) | 17.345(0.0271) | 13.121(0.0235) | 7.4735(0.0191) |
| 0.05  | 0.0476| 38.102(0.0581) | 37.095(0.0567) | 31.801(0.0565) | 25.653(0.0584) |
| 0.10  | 0.0905| 52.866(0.0762) | 50.758(0.0695) | 46.105(0.0736) | 40.855(0.0780) |
| 0.20  | 0.1633| 70.266(0.0826) | 67.536(0.0762) | 64.33(0.0828)  | 60.11(0.0878)  |
| 0.30  | 0.2201| 80.067(0.0802) | 77.598(0.0772) | 75.358(0.0844) | 70.303(0.0874) |

Table 1: Comparison of the procedures \(D\), \(D'\), and \(S\) under optimal ordered partitions.

As we see in this table, there is uniform dominance of the procedure \(S\) over \(D\) and \(D'\). For the small values of \(p\), the saving can be very significant. For example, if the mean value equals to 0.01 then the procedure \(S\) requires 13.121 tests per 100 individuals, versus 17.47 under procedure \(D'\).

**3.1.2 Lower Bound**

The connection between group testing and information theory was described by Sobel and Groll (1959) and the connection with coding theory by Sobel (1960) for the case \(p_i = p\) for all \(i\). For the comprehensive discussion, see Katona (1973).

These results can be extended to the generalized group testing problem. For example, in the case \(N = 2\) with \(q_1 > \frac{1}{2}\) and \(q_2 > \frac{1 - q_1}{q_1}\) both procedures \(D'\) and \(S\) are preferred to individual testing and optimal. The optimality follows from the fact that both procedures are equivalent to the optimal prefix Huffman code (Huffman, 1952) with the expected length \(L(N)\), \(N = 2\). For the \(N > 3\), the optimum group testing strategy does not attain \(L(N)\). It was mentioned by Sobel (1967) for the case \(p_i = p\) for all \(i\) and, therefore, it holds for the generalized group testing problem. But for the GGTP, an optimal prefix Huffman code with the expected length \(L(N)\) serves as a theoretical lower bound for the unknown optimal GGTP (Nebenzahl and Sobel, 1973).

The explicit form of \(L(N)\) is available only for the special cases with particular restrictions on the probabilities \(p_1, \ldots, p_N\) (Katona and Lee, 1972; Katona, 1973; Baranyai, 1974). It
is known that in general, the complexity of calculation of $L(N)$ is $O \left( 2^N \log_2(2^N) \right)$ due to the sorting effort. Therefore, even for small $N$, obtaining the exact value of $L(N)$ is impossible. A well-known result in information theory (Noiseless Coding Theorem; see e.g., [Katona 1973], [Cover and Thomas 2006]) provides the information theory bounds for $L(N)$:

$$H(P) \leq L(N) \leq H(P) + 1,$$

where $P = (p_1, \ldots, p_N)$ and $H(P) = \sum_{i=1}^{N} \left\{ p_i \log_2 \frac{1}{p_i} + q_i \log_2 \frac{1}{q_i} \right\}$ is the Shannon entropy. This value $H(P)$ can be used as the information lower bound for an optimal group testing procedure.

For the numerical example in Table 1, the entropy $H(P)$ was calculated for each simulation (total 1000) and mean value (standard deviation of the mean in the brackets) was reported in the last column.

4 Conclusions and Discussion

For all procedures discussed here, finding an optimal solution is equivalent to finding an optimal partition. In general, an optimum partition problem is a hard computational problem. This is why the optimal solution for the generalized group testing problem is only available for the Dorfman procedure ([Hwang 1975, 1981]). In this work we try to shed light on the additional procedures. It was shown that finding an optimal partition under procedures $S$ and $D'$ appears to be a hard computational problem. However, using an optimal ordered partition for these procedures, with the computational cost $O(N^2)$, will substantially improve the optimal procedure $D$. As a byproduct, the closed-form expression for the expected total number of tests was obtained for the Sterrett procedure. The investigation of the general class of nested procedures, to which both $D'$ and $S$ belong, is a natural next step for future work. Another direction for future investigation is to relax the assumption of error-free tests. Generally, in this case, the expected total number of tests cannot be used as the only criterion for comparison among group testing procedures and additional criteria have to be considered. [Malinovsky et al. 2016] suggested a possible criterion where procedure $D$ was investigated for the equal probabilities case. Much more
work has to be done for generalized group testing problems.

Acknowledgement

I thank James B. Orlin for the detailed explanations related to his article, and I also thank Shelemyahu Zacks for the discussions and comments on the article. The author thanks the editor, associate editor, and two referees for their time, advice, and constructive comments. The author also thanks Mattson Publishing Services for editing the article.

Appendix

A Proof of Result [1]

Proof. The proof of both parts (i) and (ii) is by induction on $k$.

(i) For $k = 2$, the result holds (Example 1). Assume that (4) is correct for $k - 1$. We prove that it is correct for $k$. Let $1_j$ be indicator function, equals 1 if the first positive individual identified is the individual $j$ ($j = 1, \ldots, n$). Let $1_0$ be the indicator function that there are no positive individuals in the group. Denote $T_k$ be the total number of the tests in the group of size $k$. It is clear that

$$T_k = T_k 1_0 + T_k 1_1 + \ldots + T_k 1_k.$$ 

We have $E(T_k 1_0) = q_1 \cdots q_k$,  $E(T_k 1_k) = q_1 \cdots q_{k-1}(1-q_k)k$,  $E(T_k 1_{k-1}) = q_1 \cdots q_{k-2}(1-q_{k-1})(k+1)$, and  $E(T_k 1_j) = q_1 \cdots q_{j-1}(1-q_j) (1 + j + E(j + 1 : k))$,  $j = 1, \ldots, k-2$, where $q_1 \cdots q_0 \equiv 1$. 

Combining all together, we have

\[ E(1:k) = E(T_k) = \sum_{j=0}^{k} E(T_{k+1}) = q_1 \cdots q_k + q_1 \cdots q_{k-1}(1 - q_k)k + q_1 \cdots q_{k-2}(1 - q_{k-1})(k + 1) \]

\[ + \sum_{j=1}^{k-2} q_1 \cdots q_{j-1}(1 - q_j)(1 + j + E(j + 1 : k)) = \]

\[- q_1 \cdots q_k (k - 1) - q_1 \cdots q_{k-1} + q_1 \cdots q_{k-2} (k + 1) + \sum_{j=1}^{k-2} q_1 \cdots q_{j-1}(1 - q_j) \]

\[ (1 + j \{ 2(k - j) - 1 - [q_{j+1} + \ldots + q_k] - [q_{j+1} \cdots q_k + \ldots + q_{k-1}q_k] \}) \]

\[ = (2k - 1) - q_1 \cdots q_k (k - 1) - q_1 \cdots q_{k-1} - q_1 \cdots q_{k-2} - q_1 \cdots q_{k-3} - \ldots - q_1q_2 - q_1 \]

\[ = (2k - 1) - q_1 \cdots q_k (k - 1) - q_1 \cdots q_{k-1} - q_1 \cdots q_{k-2} - q_1 \cdots q_{k-3} - \ldots - q_1q_2 - q_1 \]

\[ + q_1 \cdots q_k (k - 2) + q_1 \cdots q_{k-1} + q_1 \cdots q_{k-2} + q_1 \cdots q_{k-3} + \ldots + q_1q_2 \]

\[ - (q_1 + \ldots + q_{k-1}) - (q_{k-1}q_k + q_{k-2}q_{k-1}q_k + \ldots + q_2 \cdots q_{k-1}q_k) \]

\[ = (2k - 1) \]

\[ - [(q_1 + q_2 + \ldots + q_{k-1}) + q_{k-1}q_k + q_{k-2}q_{k-1}q_k + \ldots + q_2 \cdots q_k + q_1q_2 \cdots q_k]. \quad (11) \]

(ii) Recall that the people are tested in the order 1, 2, \ldots, k − 1, k with corresponding probabilities \( q_1, q_2, \ldots, q_{k-1}, q_k \). The set of these probabilities are available before we start the test process. Therefore, we are free to decide the order of the probabilities from this set, which we will assign to the persons 1, 2, \ldots, k in order to minimize the left-hand side of the equation (11). The optimal assignments to minimize the left-hand side of the equation (11) is equivalent to finding an assignment that maximizes

\[ e_{1:k} = \left(\frac{q_1 + q_2 + \ldots + q_{k-1}}{S_{1:k-1}}\right) + \left(\frac{q_{k-1}q_k + q_{k-2}q_{k-1}q_k + \ldots + q_2 \cdots q_k + q_1q_2 \cdots q_k}{R}\right). \quad (12) \]

For \( i < k \), denote \( G_{i:k} = \{q_i, \ldots, q_k\} \). We start with \( k = 2 \). We have to decide how to arrange the values of \( G_{1:2} = \{q_1, q_2\} \) in order to maximize \( q_1 + q_1q_2 \). It is obvious that we have to choose \( q_1 \geq q_2 \). Assume that \( (5) \) holds for the group of size \( k - 1 \).
with \( G_{2:k} = \{q_2, \ldots, q_k\} \), i.e.,
\[
q_{k-1} \geq \cdots \geq q_3 \geq q_2 \geq q_k,
\]
(13)
with the corresponding
\[
e_{2:k} = (q_2 + \cdots + q_{k-1}) + q_{k-1}q_k + q_{k-2}q_{k-1}q_k + \cdots + q_2 \cdots q_k.
\]
(14)
We prove that it is true for the group of size \( k \). Comparing (14) with (12) we see that \( S_{1:k-1} = q_1 + S_{2:k-1} \), \( R \) does not include \( q_1 \), and the last term \( q_1q_2\cdots q_k \) in the equation (12) is the constant for the given set of the probabilities. Therefore, to maximize (12) we must choose \( q_1 = \min_{1 \leq j \leq k-1} \{q_j\} \). Combining this with the induction step (13) we conclude that we have to decide between two alternatives \( q_{k-1} \geq \cdots \geq q_3 \geq q_2 \geq q_1 \geq q_k \) and \( q_{k-1} \geq \cdots \geq q_3 \geq q_2 \geq q_k \geq q_1 \). Direct calculation shows that the first one will maximize (12) (or, as an alternative minimize (1)).

\[\square\]

B Dynamic Programming Algorithm for Finding an Optimal Ordered Partition

We need some additional notation to present this well-known algorithm (see, e.g., Hwang (1975); Chakravarty et al. (1982)). Let \( U_N = \{u_1, u_2, \ldots, u_N\} \) be the population subject to classification. Each unit \( u_i \) has the probability \( p_i \) to be defective and the probability \( q_i = 1 - p_i \) to be good. Without loss of generality, assume that the units \( u_1, \ldots, u_N \) are labeled such that \( 0 < p_1 \leq p_2 \leq \ldots \leq p_N < 1 \). The goal is to find an optimal ordered partition under procedure \( A \), which is a solution of optimization problem (1) under ordered partition restriction. Denote the cost function \( C_A(U_N) \) as an expected total number of tests for the population \( U_N \) under an optimum ordered partition using procedure \( A \). DP algorithm solution is
\[
C_A(U_0) = 0, \quad C_A(U_1) = 1, \quad C_A(U_k) = \min_{0 \leq i \leq k-1} \{E_A(U_k - U_i) + C_A(U_i)\}, \quad 2 \leq k \leq N,
\]
(15)
where \( E_A(U_k) \) is the expected total number of tests of testing group \( U_k \) under procedure \( A \), \( U_0 = \emptyset \).

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