Validation of a non-invasive method for the early detection of metabolic syndrome: a diagnostic accuracy test in a working population

Manuel Romero-Saldaña,1 Pedro Tauler,2,3 Manuel Vaquero-Abellán,4 Angel-Arturo López-González,5 Francisco-José Fuentes-Jiménez,6,7 Antoni Aguiló,3,8 Carlos Álvarez-Fernández,1 Guillermo Molina-Recio,9 Miquel Bennasar-Veny2,8

ABSTRACT

Objectives A non-invasive method for the early detection of metabolic syndrome (NIM-MetS) using only waist-to-height ratio (WHR) and blood pressure (BP) has recently been published, with fixed cut-off values for gender and age. The aim of this study was to validate this method in a large sample of Spanish workers.

Design A diagnostic test accuracy to assess the validity of the method was performed.

Setting Occupational health services.

Participants The studies were conducted in 2012–2016 on a sample of 60 799 workers from the Balearic Islands (Spain).

Interventions The NCEP-ATP III criteria were used as the gold standard. NIM-MetS has been devised using classification trees (the χ² automatic interaction detection method).

Main outcome measures Anthropometric and biochemical variables to diagnose MetS. Sensitivity, specificity, validity index and Youden Index were determined to analyse the accuracy of the diagnostic test (NIM-MetS).

Results With regard to the validation of the method, sensitivity was 54.7%, specificity 94.9% and the Validity Index 91.2%. The cut-off value for WHR was 0.54, ranging from 0.51 (lower age group) to 0.56 (higher age group). Variables more closely associated with MetS were WHR (area under the curve (AUC)=0.85; 95% CI 0.84 to 0.86) and systolic BP (AUC=0.79; 95% CI 0.78 to 0.80)). The final cut-off values for the non-invasive method were WHR ≥0.56 and BP ≥128/80 mm Hg, which includes four levels of MetS risk (very low, low, moderate and high).

Conclusions The analysed method has shown a high validity index (higher than 91%) for the early detection of MetS. It is a non-invasive method that is easy to apply and interpret in any healthcare setting. This method provides a scale of MetS risk which allows more accurate detection and more effective intervention.

INTRODUCTION

The obesity epidemic which currently affects the world population has resulted in a general increase in the prevalence of metabolic syndrome (MetS).1-8 Overweight and obesity are factors related to the onset of type 2 diabetes, hypertension, dyslipidaemia and cardiovascular diseases (CVDs). In particular, central obesity, which is defined as an excessive accumulation of abdominal fat, is an important predictor of cardiovascular risk and MetS.4 5 Metabolic syndrome is defined as a pluripathological state characterised by the joint presence of several cardiovascular risk factors such as abdominal obesity, high blood pressure (BP), altered glucose and lipid metabolism (low high-density lipoprotein (HDL)-cholesterol and high triglycerides).6

Although there are several analytical/instrumental techniques for measuring the amount and distribution of body fat, there is no consensus on the ideal method to calculate central adiposity, nor on how to decide which cut-off points provide greater accuracy, efficiency, sensitivity (S) and specificity (SP) in all cases.7 8
A simple and inexpensive alternative to these instruments as a way of quantifying abdominal fat is to make anthropometric measurements of central obesity.\(^9\) Waist circumference (WC), body mass index (BMI), waist-to-height ratio (WHtR), waist-to-hip ratio, hip-to-height ratio, Body Adiposity Index (BAI), Visceral Adiposity Index, A Body Shape Index (ABSI) and percentage of body fat (\(\%BF\)) are some examples that can be found in numerous epidemiological studies, in which they try to indirectly relate intra-abdominal (visceral) fat with parameters such as morbidity and mortality, and also with prevalence of hypertension, diabetes, MetS, and so on.\(^{10-13}\)

Since the mid-1990s, WHtR has been the most widely used anthropometric indicator and the one which has obtained the best predictive results for cardiovascular risk.\(^{14}\) In a previous publication, a non-invasive method for early detection of MetS (NIM-Mets) using only two anthropometric variables (WHtR and BP) has been proposed and validated.\(^{15}\) This method suggests WHtR ≥0.55 as the predictive threshold for the early detection of MetS for both men and women and, also, for any age stratum.

The aim of this study is to validate the NIM-Mets method in a large representative sample of Spanish workers, to determine its predictive ability and to find out the stability of the cut-off value of WHtR ≥0.55 by gender and age.

**MATERIAL AND METHODS**

**Design and sample**

A diagnostic test using a cross-sectional study was carried out on a working population from the Balearic Islands (Spain) between 2012 and 2016. Subjects participating in the study were randomly selected during their work health periodic assessments. Every day, each worker was assigned a number and half of the examined workers were randomly selected using a random number table. A total of 69,581 workers were invited to participate in the study. However, 8,782 (12.6\%) refused to participate and, thus, the final number of participants was 60,799 workers (10.2\% of the active population) belonging to different economic sectors (public administration, health services, etc), aged 20–70 years; 57.3\% of the participants was male 42.7\% female.

Participants were informed of the purpose of the study before they provided written informed consent to participate. The study protocol complied with the Declaration of Helsinki for conducting medical research involving human subjects.

**Data collection and definition of variables**

To carry out the anthropometric measurements, recommendations contained in the manual ‘International Standards for Anthropometric Assessment (ISAK)’ were followed.\(^{16}\) All the measurements were made by specifically trained staff in order to minimise the variation coefficients. Each measurement was performed three times, taking the average as the final value.

The independent variables were classified into the following categories:

- **Personal and health habits**: gender, age and tobacco consumption.
- **Anthropometric measurements**:
  - WC in cm.
  - BMI, calculated as body weight (kg) divided by height (m) squared, in kg/m\(^2\).
  - \(\%BF\) calculated according to the Deurenberg equation: \(\%BF = 1.2 \times (\text{BMI}) + 0.23 \times (\text{age in years}) - 10.8 \times (\text{gender}) - 5.4.\) Gender: female (0), male (1)
  - Waist-to-height ratio (WHtR), calculated as WC divided by height, both in cm.
  - ABSI calculated as WC/((\text{BMI})^{2/3}(\text{height})^{1/2}).
- **Blood measurements**:
  - Systolic BP (SBP) in mm Hg.
  - Diastolic BP (DBP) in mm Hg.
  - Total cholesterol (mg/dL), low-density lipoprotein; LDL-cholesterol (mg/dL), HDL-cholesterol (mg/dL), glucose (mg/dL) and triglycerides (mg/dL).

Body weight was measured to the nearest 0.1 kg with an electronic scale (Seca 700 scale, Seca GmbH, Hamburg). Height was measured to the nearest 0.5 cm with a stadiometer (Seca 220 (CM) telescopic height rod for column scales, Seca GmbH, Hamburg). WC was measured halfway between the lower costal border and the iliac crest. The measurement was taken at the end of a normal expiration with the subject standing up, with their feet together and their arms hanging down by their sides.

Venous blood samples were taken from the antecubital vein in suitable vacutainers without anticoagulant in order to obtain serum. The blood samples were taken after a 12-hour overnight fast. Participants sat and rested for at least 15 min before the blood samples were taken. Serum was obtained after centrifugation (15 min, 1000 ×g, 4°C) of the blood samples. The serum was stored at −20°C and analyses were performed within 3 days. Concentrations of glucose, cholesterol and triglycerides were measured in serum following the standard procedures used in clinical biochemistry laboratories with an autoanalyser (SYNCHRON CXH9 PRO, Beckman Coulter, Brea, California, USA).

BP was determined after the subjects had rested in the supine position for 10 min, using an automatic and calibrated sphygmomanometer (OMRON M3, OMRON Healthcare Europe, Spain). As in the case of the anthropometrical measurements, BP was measured three times, leaving a 1 min gap between each measurement, and the average value was then calculated.

Presence of MetS was ascertained by using the criterion suggested by the NCEP-ATP III definition (when three of five of the following characteristics are present, a diagnosis of metabolic syndrome can be made):

- **Abdominal obesity** (WC ≥102 cm in men and WC ≥88 cm in women).
Triglycerides $\geq 150$ mg/dL.
- HDL-cholesterol $< 40$ mg/dL in men and $< 50$ mg/dL in women.
- BP $\geq 130/85$ mm Hg.
- Fasting glucose $\geq 100$ mg/dL.

**Non-invasive method for early detection of MetS**

NIM-Mets is a new tool for screening for MetS based on the following anthropometric variables and cut-off values: WHtR $\geq 0.55$ and BP $\geq 128/85$ mm Hg.

This method classifies the population into two groups with different levels of risk:

- **Workers with high risk of MetS (probability $> 61.7\%$):** this group would contain those subjects with both positive variables, that is, WHtR $\geq 0.55$ and BP $\geq 128/85$ mm Hg.
- **Workers with low risk of MetS (probability of 0.5\%—16.9\%):** this group would contain those subjects who have any of the other possible combinations between the two variables considered.

**Statistical analysis**

Statistical analysis was carried out using IBM SPSS Statistics V.22.0 software (SPSS, IBM, Chicago, Illinois, USA) and EpiData V.4.2. (Department of Sanidade, Xunta de Galicia, Galicia, Spain). Continuous data are presented as mean values, SD and 95% CI. Categorical data are shown as frequency counts and percentages. All the data were tested for their normal distribution (Kolmogorov–Smirnov test with Lilliefors adjustment).

Student’s t-test and Mann-Whitney U test were used in the bivariate analyses for normal and non-normal distributed variables, respectively. Analysis of variance tests with the post hoc Bonferroni contrast method were carried out when more than two groups were considered in the analysis. The Levene test was used to determine variance equality. The $\chi^2$ test was applied to assess differences between groups in categorical variables.

Receiver operating characteristic (ROC) curves were performed and the area under the curve (AUC) was calculated to find which explanatory variables best predict the onset of MetS. We obtained the cut-off value for each explanatory variable through the Youden Index (JI) as $JI = S + SP - 1$.

To measure the accuracy of the diagnostic test, S, SP, positive and negative predictive values (PPV and NPV), likelihood ratios (LH+ and LH−), Validity Index (VI) and JI were analysed. VI was calculated as the quotient between the sum of true positives and true negatives, divided by the total number of subjects, therefore representing the percentage of subjects properly classified by the test.

The modification of NIM-Mets was obtained from a clinical decision tree (classification) using the $\chi^2$ automatic interaction detection (CHIAD) technique as a growth method. The statistical significance level for splitting nodes and merging categories was $p < 0.05$, and significance values were corrected by the Bonferroni method, with a maximum number of iterations of 2000.

The level of statistical significance was fixed in all the contrasts for an $\alpha$ error below 5%, and the CIs were calculated with a 95% level of confidence.

**Patient and public involvement**

Patients were not involved in setting the research question and in the study design. All patients were randomly selected during their work health periodic assessments to participate in the study and they were interviewed face to face by trained researchers for detailed explanation of the purpose of this research and informed consent at the beginning. No patients were involved in data analysis or manuscript writing. Results of the research will not be disseminated to the patients.

**RESULTS**

**Characteristics of the study sample**

Of the 60 799 workers, 34 827 were male (57.3%). The overall mean age was 40 years (39.9–40.1) (table 1). Among anthropometrical and blood parameters shown in table 1, women showed higher %BF and HDL-cholesterol values ($p < 0.001$), while men showed significantly higher values for the rest of the parameters shown in this table. The prevalence of smokers was 34.8% (36.6% in men and 32.5% in women) and 17.6% of participants were obese (20.0% men and 14.4% women). As regards to drug treatments, 6.6% of participants were undergoing antihypertensive treatment, 3.2% lipid-lowering treatments and 1.7% antidiabetic treatments. Finally, the overall prevalence of MetS was 9.0%, with 11.8% in men and 5.4% in women ($p < 0.001$).

**NIM-Mets validation**

Table 2 shows the results of diagnostic tests after applying the NIM-Mets method compared with NCEP-ATP III as a control test. Overall, the indicators of the NIM-Mets method validation were as follows: $S = 54.7\%$ (53.4–56.0), SP$ = 94.9\%$ (94.7–95.0) and VI$ = 91.2\%$ (91.0–91.5). S was higher in men (59.4\%) than in women (40.9\%).

As regards the NIM-Mets safety indicators, results in the total sample were: PPV$ = 51.3\%$ (50.0–52.6) and NPV$ = 95.5$ (95.3–95.7). By gender, PPV was higher in men (51.4\%) than in women (50.8\%), while NPV was higher in women (96.7\%) than in men (94.5\%). Finally, the overall JI was 0.50 (0.48 to 0.51), higher in men (0.52) than in women (0.39).

A second question to be dealt with in this research was to compare the cut-off value for WHtR proposed by NIM-Mets with that obtained in the study sample, and thus determine its variability according to the gender variable and in different age groups (table 3). In the whole sample ($n = 60799$), a cut-off value of 0.54 was obtained for WHtR. In the group of men ($n = 34827$) the resulting threshold was 0.56, while for women ($n = 25972$) it was 0.53.

It can be seen how the cut-off point increases with age. For men, it ranged from 0.55 (age 20–30 years) to 0.56
(age ≥ 51 years), whereas for women it was between 0.51 (age 20–30 years) and 0.55 (age ≥ 51 years). The differences between the cut-off values for men and women become narrower as the age increased.

Figure 1 shows the results for the anthropometric variables’ ROC curves. WHtR achieved the highest AUC 0.85 (95% CI 0.84 to 0.86), with a cut-off value of 0.54, reaching top values of S=68.5%, SP=87.0% and JI=0.56. The second variable with the highest AUC was WC, with 0.83 (95% CI 0.82 to 0.84), a cut-off value of 89.1 cm and S=72.5, SP=77.6% and JI=0.5. BMI with an AUC=0.80 and SBP with AUC=0.79 also stood out.

After that, different clinical decision trees were made with a range of cut-off values for WHtR and BP (table 4).

### Table 1 Characteristics of the sample according to gender

| Variable       | Total n=60 799 | Men n=34 827 | Women n=25 972 | P values |
|----------------|----------------|--------------|----------------|----------|
| Age (years)    | 40.0 (39.9 to 40.1) | 40.4 (40.3 to 40.5) | 39.5 (39.3 to 39.6) | <0.001   |
| Smoker (yes)   | 21 177 (34.8%)  | 12 746 (36.6%)  | 8431 (32.5%)  | <0.001   |
| BMI (kg/m²)    | 26.1 (26 to 26.1)  | 26.9 (26.8 to 26.9)  | 25.0 (25 to 25.1)  | <0.001   |
| WC (cm)        | 82.9 (82.9 to 83.0) | 88.6 (88.5 to 88.7) | 75.4 (75.3 to 75.5) | <0.001   |
| WHtR           | 0.49 (0.49 to 0.49) | 0.51 (0.50 to 0.51) | 0.47 (0.46 to 0.47) | <0.001   |
| ABSI           | 0.07 (0.07 to 0.07) | 0.07 (0.07 to 0.07) | 0.07 (0.07 to 0.07) | <0.001   |
| %BF            | 28.9 (28.9 to 29.0) | 25.3 (25.3 to 25.4) | 33.7 (33.6 to 37.8) | <0.001   |
| SBP (mm Hg)    | 120.8 (120.6 to 120.9) | 125.4 (125.2 to 125.6) | 114.6 (114.4 to 114.8) | <0.001   |
| DBP (mm Hg)    | 73.6 (73.5 to 73.7) | 76.0 (75.9 to 76.1) | 70.4 (70.3 to 70.5) | <0.001   |
| Glucose (mg/dL)| 88.3 (88.1 to 88.5) | 90.6 (90.4 to 90.8) | 85.2 (85 to 85.4) | <0.001   |
| Cholesterol (mg/dL) | 195.2 (194.9 to 195.5) | 196.9 (196.5 to 197.3) | 193.0 (192.6 to 193.4) | <0.001   |
| HDL-cholesterol (mg/dL) | 52.4 (52.3 to 52.5) | 50.5 (50.4 to 50.6) | 55.0 (54.9 to 55.1) | <0.001   |
| LDL-cholesterol (mg/dL) | 121.2 (120.9 to 121.5) | 121.8 (121.4 to 122.6) | 120.5 (120.2 to 120.9) | <0.001   |
| Triglycerides (mg/dL) | 109.3 (108.7 to 109.9) | 125.3 (124.4 to 126.2) | 88.8 (88.2 to 89.4) | <0.001   |
| MetS (yes)     | 5587 (9.0%)  | 4097 (11.8%)  | 1390 (5.4%)  | <0.001   |

A value of p<0.05 indicates statistically significant differences between genders.

ABSI, A Body Shape Index; %BF, percentage of body fat calculated according to the Deurenberg equation; BMI, body mass index; DBP, diastolic blood pressure; HDL, high-density lipoprotein; LDL, low-density lipoprotein; MetS, metabolic syndrome; SBP, systolic blood pressure; WC, waist circumference; WHtR, waist-to-height ratio.

### Table 2 Diagnostic test accuracy of the non-invasive method for early detection of metabolic syndrome against NCEP-ATP III

|                | Total | Men | Women |
|----------------|-------|-----|-------|
|                | Yes   | No  | Total | Yes   | No  | Total | Yes   | No  | Total |
| NIM-MetS (n)   |       |     |       |       |     |       |       |     |       |
| Yes            | 3001  | 2850| 5851  | 2433  | 2300| 4733  | 568   | 550 | 1118  |
| No             | 2486  | 5246| 54948 | 1664  | 28430| 30094 | 822   | 24032| 24854 |
| Total          | 5487  | 55312| 60799 | 4097  | 30730| 34827 | 1390  | 24582| 25972 |

Efficacy indicators, 95% CI

|                | Total |
|----------------|-------|
| Sensitivity    | 54.7 (53.4 to 56.0) |
| Specificity    | 94.9 (94.7 to 95.0) |
| PPV            | 51.3 (50.0 to 52.6) |
| NPV            | 95.5 (95.0 to 95.7) |
| VI             | 91.2 (91.0 to 91.5) |
| LH +           | 10.6 (10.2 to 11.1) |
| LH -           | 0.48 (0.46 to 0.49) |
| JI             | 0.50 (0.48 to 0.51) |

Jl, Youden Index; LH+, positive likelihood ratio; LH−, negative likelihood ratio; NIM-MetS, non-invasive method for early detection of metabolic syndrome; NPV, negative predictive value; PPV, positive predictive value; VI, Validity Index.
Thus, the range of cut-offs for WHtR was defined by eight thresholds between 0.54 and 0.57, and included, among others, the cut-off value proposed by NIM-MetS (WHtR ≥ 0.55) and the cut-off value for the total sample (WHtR ≥ 0.54). In addition, three models were established for BP: BP ≥ 128/80 mm Hg (cut-off values obtained for SBP and DBP as ROC curves, shown in figure 2); BP ≥ 128/85 mm Hg (BP cut-off values proposed by NIM-MetS); and finally, only SBP ≥ 128 mm Hg (second covariate with the highest adjusted OR in the multiple logistic regression). In this way, 24 clinical decision trees were set up using CHAID methodology. Each tree comprised of a parent node (node 0), two primary subsidiary nodes (nodes 1 and 2) and four secondary subsidiary nodes (nodes 3, 4, 5 and 6). Each of the last four nodes denoted the probability of having MetS. Thus, node 3 corresponds to the probability that a worker has MetS when both anthropometric variables are negative (below cut-off values). Node 4 indicates the probability that a worker has MetS when BP is above the cut-off value and WHtR is below. Node 5 represents the probability that a worker has MetS when BP is lower than the cut-off value and WHtR is above. Finally, node 6 shows the probability that a worker suffers from MetS when both variables are positive (above the cut-off values). The model BP ≥ 128/80 mm Hg was chosen because it had the highest JI value (greatest S and SP combined) at each of the WHtR cut-off points (table 4).

The next step was to select the final cut-off value for WHtR. To do this, the method’s probability of detection (node 6 value) and the JI for the BP model chosen (BP ≥ 128/80 mm Hg) were plotted for each WHtR cut-off point (figure 2).

It was noted that the probability of detection of MetS in each tree (node 6 value) and the WHtR threshold, as well as the JI and the WHtR threshold, follow linear functions, in which the equations of its lines are as follows:

- Probability MetS (node 6) = 5.534*WHtR−2.58
- JI = −1.758*WHtR +1.486

Thus, the final threshold value for WHtR was determined by the cut-off points of both lines: WHtR=0.56 (figure 2).

The resulting new method for the early detection of MetS (new NIM-MetS) includes these conditions: WHtR ≥ 0.56 and BP ≥ 128/80 mm Hg. Figure 3 shows the decision tree created from these variables and cut-off points.

The S of the proposed method was 56.4%, SP was 94.5%, VI was 91.1% and JI was 0.51.

Finally, from the probabilities obtained in nodes 3, 4, 5 and 6, a risk gradient for MetS was developed, according to the WHtR and BP covariates and the proposed cut-off values. Thus, those subjects with lower WHtR and BP values than the cut-off point have a very low probability of suffering from MetS (PMetS=0.4%). Low risk (PMetS=8.3%) would be found only in those individuals with BP values over 128/80 mm Hg, but a low WHtR. A moderate level of risk (PMetS=16.3%) would include normotensive subjects who had a WHtR ≥ 0.55. Finally,

Table 3  Area under the curve (AUC) and cut-off values for waist-to-height ratio (WHtR) according to gender and age groups

| Age group (years) | n     | Prevalence of MetS (%) | AUC 95% CI | Cut-off value | Sensitivity (%) | Specificity (%) | JI    |
|------------------|-------|------------------------|------------|--------------|----------------|----------------|-------|
| **Men**          |       |                        |            |              |                |                |       |
| 20–30            | 6825  | 3.1                    | 0.92 (0.9 to 0.95) | 0.55          | 80.3           | 93.4           | 0.74  |
| 31–40            | 11623 | 7.5                    | 0.88 (0.86 to 0.89) | 0.55          | 77.4           | 88.1           | 0.65  |
| 41–50            | 10080 | 14.9                   | 0.82 (0.81 to 0.83) | 0.56          | 66.4           | 87.7           | 0.54  |
| ≥51              | 6659  | 23.1                   | 0.75 (0.74 to 0.77) | 0.56          | 58.9           | 83.0           | 0.42  |
| **Total**        | 34827 | 11.8                   | 0.84 (0.83 to 0.85) | 0.56          | 66.7           | 88.8           | 0.56  |
| **Women**        |       |                        |            |              |                |                |       |
| 20–30            | 5715  | 1.1                    | 0.90 (0.85 to 0.95) | 0.51          | 82.0           | 84.0           | 0.74  |
| 31–40            | 8529  | 2.7                    | 0.91 (0.89 to 0.93) | 0.53          | 80.3           | 90.8           | 0.71  |
| 41–50            | 7641  | 6.6                    | 0.91 (0.89 to 0.93) | 0.53          | 80.3           | 90.8           | 0.71  |
| ≥51              | 4087  | 14.4                   | 0.75 (0.73 to 0.77) | 0.55          | 48.4           | 90.5           | 0.39  |
| **Total**        | 25972 | 5.4                    | 0.85 (0.84 to 0.86) | 0.53          | 65.1           | 88.7           | 0.54  |
| **Total**        |       |                        |            |              |                |                |       |
| 20–30            | 12540 | 2.1                    | 0.92 (0.9 to 0.94) | 0.53          | 84.4           | 90.1           | 0.75  |
| 31–40            | 19792 | 5.5                    | 0.90 (0.89 to 0.91) | 0.54          | 78.2           | 88.5           | 0.67  |
| 41–50            | 17721 | 11.3                   | 0.83 (0.82 to 0.84) | 0.54          | 69.6           | 84.3           | 0.54  |
| ≥51              | 10746 | 19.8                   | 0.76 (0.75 to 0.77) | 0.56          | 57.0           | 85.3           | 0.42  |
| **Total**        | 60799 | 9.0                    | 0.85 (0.84 to 0.86) | 0.54          | 68.5           | 87.0           | 0.56  |

*MetS according to NCEP ATP III criterion.
JI, Youden Index; MetS, metabolic syndrome.
subjects with WHtR ≥ 0.56 and BP ≥ 128/80 mm Hg, would have a 50.5% risk of having MetS.

**DISCUSSION**

The validation of NIM-MetS was carried out from a study of diagnostic tests conducted in Spanish Caucasian adult workers and using, as a reference test, the NCEP-ATP III criteria for the diagnosis of MetS. The early detection of MetS is the key to improving the quality of life of our population, since it prevents the appearance of associated complications such as CVD, type 2 diabetes mellitus and, even, cancer.17–20

NIM-MetS has shown high VI in men (88.6%) and women (94.7%). Overall, for every 100 workers, the method properly classified 91 cases. Similarly, NIM-MetS has proved to be highly specific, reaching an overall SP of 94.9% (92.5% in men and 97.8% in women). Both VI and the SP recorded values above those achieved by this method in another Spanish population,15 where it obtained an IV = 89.5% and 91.5% SP. For S, the overall figure was 56.4% (59.4% for men and 40.9% for women), while in the original population, the overall S was 77.9%. Because it supposes a simple, easy to apply even in large populations and non-invasive method, it could be defined as a useful method in spite of the S found in the present study could be considered as moderate. The high SP together with the high VI shown for the screening of the cardiometabolic risk are characteristics that increase the acceptability of the method.

Although the indicators of validation and accuracy of a screening test (S and SP) are intrinsic properties of the test itself and do not depend on the prevalence of the disease considered, this does not prevent these indicators from being affected by characteristics of the population they are applied to.21 In fact, the most common observation is that a test for early detection or diagnosis alters its S and SP depending on these characteristic features of the population. Therefore, the main differences between the two populations (the Balearic and the one considered in the previous study developed in Cordoba15) were
analysed, highlighting those features of the Balearic population which contributed to a decreased S: a younger population (40.0 years vs 45.1 years), more women (42.7% vs 32.1%), more smokers (34.8% vs 28.6%) and lower values for WC (82.9 cm vs 87.8 cm), WHtR (0.49 vs 0.52) and BMI (26.1 kg/m² vs 26.5 kg/m²).

As regards the safety indicators of the test, the PPV and NPV, they are definitely affected by the prevalence of the population, lowering the PPV when the prevalence of MetS is lower. In this way, although we found a lower prevalence of MetS in the Balearic Islands than in Cordoba (9.0% vs 13.9%), the NIM-Mets produced a lower PPV in the Balearic Islands (51.3% vs 61.7%), while the NPV remained very similar (95.5% vs 95.9%).

Screening tests are often used in clinical practice. However, there are very few methods for the early detection of MetS other than the currently known diagnostic criteria, and there are even fewer non-invasive screening tests. A study in the Republic of Korea examined the validity of a test for the early detection of MetS based on the muscle-to-fat ratio.22 The study was conducted on 6256 participants, with an S of 68.6% in men and 76.0% in women, and an SP of 63.8% in men and 53.8% in women. Miller et al23 proposed a screening method for MetS in 745 young adults (18–29 years old) in USA, based on making decision trees with the CHAID methodology and using all the criteria for MetS. The method had a validity rate of 89.4% and an S rate of 61.7%. Finally, De Kroon et al24 conducted a screening test for MetS in 642 young people (aged 17–28 years) in the Netherlands using anthropometric variables (BMI, WC and BP). The S of the method was 68.7% and the VI was 95.6%.

Table 4

| WHtR range | BP model | Probabilities of MetS for nodes in the decision tree | Efficacy indicators for diagnostic test accuracy |
|------------|----------|------------------------------------------------------|-------------------------------------------------|
|            |          | Sensitivity (%) | Specificity (%) | VI (%) | JI     |
| Cut-off points | BP  | Node 3 | Node 4 | Node 5 | Node 6 |       |       |       |
| 0.535      | BP*    | 0.3    | 7.4   | 10.3  | 38.4  | 63.9  | 89.7  | 87.5  | 0.54  |
|           | BP†    | 0.5    | 10.4  | 12.2  | 43.3  | 59.3  | 92.3  | 89.3  | 0.52  |
|           | SBP    | 0.6    | 10.5  | 13.7  | 43.3  | 57.0  | 92.6  | 89.4  | 0.5   |
| 0.540      | BP*    | 0.3    | 7.6   | 11.5  | 40.7  | 62.7  | 90.9  | 88.4  | 0.54  |
|           | BP†    | 0.5    | 10.5  | 13.5  | 45.7  | 57.8  | 93.2  | 90.0  | 0.51  |
|           | SBP    | 0.6    | 10.7  | 15.2  | 45.7  | 55.6  | 93.5  | 90.0  | 0.49  |
| 0.544†     | BP*    | 0.4    | 7.7   | 12.5  | 42.8  | 61.6  | 91.8  | 89.1  | 0.53  |
|           | BP†    | 0.5    | 10.9  | 14.6  | 48.0  | 56.7  | 93.9  | 90.5  | 0.51  |
|           | SBP    | 0.6    | 10.8  | 16.4  | 48.0  | 54.5  | 94.1  | 90.6  | 0.49  |
| 0.550‡     | BP*    | 0.4    | 7.9   | 14.2  | 46.1  | 59.6  | 93.1  | 90.1  | 0.53  |
|           | BP†    | 0.5    | 10.9  | 16.6  | 51.3  | 54.7  | 94.8  | 91.2  | 0.46  |
|           | SBP    | 0.7    | 11.0  | 18.5  | 51.4  | 52.6  | 95.1  | 91.2  | 0.48  |
| 0.555      | BP*    | 0.4    | 8.2   | 15.7  | 49.1  | 57.8  | 94.1  | 90.8  | 0.52  |
|           | BP†    | 0.6    | 11.1  | 18.3  | 54.4  | 53.0  | 95.6  | 91.7  | 0.49  |
|           | SBP    | 0.7    | 11.3  | 20.3  | 54.5  | 51.0  | 95.8  | 91.7  | 0.47  |
| 0.560      | BP*    | 0.5    | 8.5   | 17.0  | 51.8  | 55.7  | 94.9  | 91.3  | 0.51  |
|           | BP†    | 0.6    | 11.5  | 19.9  | 57.1  | 51.0  | 96.2  | 92.1  | 0.47  |
|           | SBP    | 0.8    | 11.6  | 21.9  | 57.2  | 49.1  | 96.4  | 92.1  | 0.46  |
| 0.565      | BP*    | 0.5    | 8.8   | 18.6  | 54.9  | 53.4  | 95.6  | 91.8  | 0.49  |
|           | BP†    | 0.6    | 11.9  | 21.8  | 60.3  | 48.8  | 96.8  | 92.5  | 0.46  |
|           | SBP    | 0.8    | 12.0  | 23.9  | 60.4  | 47.0  | 96.9  | 92.4  | 0.44  |
| 0.570      | BP*    | 0.5    | 9.1   | 19.9  | 57.4  | 51.4  | 96.2  | 92.2  | 0.48  |
|           | BP†    | 0.7    | 12.3  | 23.3  | 62.8  | 46.9  | 97.2  | 92.7  | 0.44  |
|           | SBP    | 0.9    | 12.4  | 25.5  | 63.0  | 45.2  | 97.4  | 92.7  | 0.43  |

*BPH ≥128/80 mm Hg; †BP ≥128/85 mm Hg; ‡Cut-off point for WHtR obtained in the total simple (n=60 799); §Cut-off point proposed by NIM-MetS.

BP, blood pressure; JI, Youden Index; NIM-MetS, non-invasive method for early detection of metabolic syndrome; SBP, systolic blood pressure ≥128 mm Hg; VI, Validity Index; WHtR, waist-to-height ratio.
Another hypothesis put forward in this research was to test whether the cut-off value proposed by NIM-MetS for WHtR ($\geq 0.55$) would be reproduced in a large sample (60,799 workers from the Balearic Islands), and if it was also valid for both men and women and also for different age groups. WHtR had a cut-off value of 0.54 for the total sample, with 0.56 men and 0.53 for women. As regards age groups, the WHtR threshold increased with age, with 0.55–0.56 for men and 0.51–0.55 for women. These differences were reduced in the total sample (0.53–0.56).

It is noteworthy that several authors have proposed a universal 0.50 cut-off point for WHtR, both to detect MetS and to predict cardiometabolic risk and overall cardiovascular mortality. However, in Spain, a cross-sectional

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**Figure 2** Waist-to-height ratio (WHtR) cut-off point resolution. MetS, metabolic syndrome.

**Figure 3** Definitive decision tree, new NIM-Mets proposed. BP, blood pressure, NIM-MetS, non-invasive method for early detection of metabolic syndrome; WHtR, waist-to-height ratio.
study in the general population (n=6279, mean age=43 years) showed that WHtR was the best anthropometric predictor of MetS (NCEP-ATP III), and the authors proposed a cut-off value of 0.55, with which they obtained an S of 91.0% and an SP of 64.0%. This cut-off value is very similar to the one proposed from the results obtained in the present study.

In Chile, two important child population studies were conducted by Arnaiz et al, showing results which match the value of the cut-off point proposed in the present study. Thus, in the first study, conducted on 209 schoolchildren (mean age of 11.5 years), the authors obtained a cut-off value of 0.55 WHtR for MetS, while in the second study, performed in 2980 children aged 6–14 years (mean age of 9.9), the authors concluded that the WHtR did not change with age and gender and, therefore, a universal cut-off value could be agreed on for both children and adults.

The prospective study by Koch et al carried out in Chile on about 6714 men and 6340 women, evaluating the relationship between various anthropometric indices of adiposity, cardiovascular risk factors and mortality for a cut-off value of 0.55 obtained an S of 75.8% and SP of 73.3% for men, and an S of 77.6% and SP of 56.3% for women.

In addition, several investigations conducted in non-European and non-Hispanic populations also concur on this threshold of WHtR for MetS. Thus, Obeidat et al in a study on a Jordanian population (n=630, aged 20–70 years) reported a cut-off value of 0.56 in men and 0.52 in women; in India, Rajput et al achieved a threshold of 0.52 for men and women (n=3042) in all locations (rural or urban areas); and in China, He et al in a descriptive study of 1068 adult subjects, reported a cut-off value for WHtR of 0.5 for men and women alike.

When the NIM-MetS method was applied in this new larger sample of 60 799 workers from the Balearic Islands, the method has showed again the same variables obtained in the original study performed in Cordoba. In the multiple logistic regression, WHtR and BP achieved the highest adjusted OR values. Thus, WHtR was the anthropometric index that best discriminated MetS presence, with an adjusted OR value of 4.4 (3.9–4.9), while SBP obtained an adjusted OR value of 3.8 (3.5–4.1). In addition, the cut-off values obtained for WHtR and for BP are very similar to those of the original method.

Several investigations have confirmed the high predictive ability of WHtR for MetS and CVD. In the systematic review conducted by Ashwell et al, in which 10 out of the 31 studies analysed the association between anthropometric measurements of central obesity and MetS, WHtR had the highest AUC value of 0.76 (men) and 0.75 (women). This contrasted with WC, which obtained an AUC value of 0.75 (equal for men and women) and BMI, with an AUC value of 0.72 (men and women). Similarly, a meta-analysis conducted by Savva et al, in which 8 out of the 24 studies included compared WHtR (cut-off point 0.5) with BMI (cut-off point of 23 for the Asian population and 25 for the rest) for cardiometabolic risk in Asian and non-Asian populations, concluded that WHtR showed a stronger association with MetS than with BMI.

Through the present study, the NIM-MetS method has been reproduced, and definitive cut-off values have been proposed for WHtR (0.56) and BP (128–80 mm Hg), from which an S rate of 56.4%, an SP rate of 94.5%, a VI of 91.1% and a JI of 0.51 are obtained. On the other hand, finally, the long-term ability of MetS to predict CVD has been shown to be limited by the dichotomous (binary) and qualitative nature of the classic diagnostic criteria for MetS. An innovative aspect that NIM-MetS brings is to provide a gradient or scale of risk of developing MetS which is divided into four risk levels: very low risk (p=0.4%), low risk (p=8.3%), moderate risk (p=16.4%) and high risk (p=50.5%). In this way, health professionals can take certain steps depending on the level of risk of MetS and promote a more accurate and early detection of the possible complications associated with CVD and MetS. Along the same lines, there have been several studies using methods based on scores to quantify the amount of risk accumulated by the presence of the components that define the MetS (Metabolic Syndrome Severity Score).

Limitations
This study presents some limitations that must be acknowledged. First, we must bear in mind that there are different definitions and criteria to determine the presence of MetS. In this study, the presence of MetS was ascertained using the NCEP-ATP III definition as the gold standard, which supposes one of the definitions most used and widely accepted by the international community and the WHO. In addition, the study data refer to the Caucasian population. Thus, results could not have great applicability to other populations.

Although in the present study NIM-MetS methodology has been tested in a very large sample of workers, the S found was lower than that obtained in the original study leading to the proposal of the method. This could be related to differences in the study samples, with the workers from the Balearic Islands showing lower prevalence of SMet and obesity and being younger. Although the prevalence of MetS does not affect S and SP, this lower prevalence influences PPV and NPV.

Although the percentage of participation is high (87.4%), we should take into account that it is not the total target population and, therefore, a bias could have been introduced in the results. Furthermore, participants highly concerned about their health, and thus probably healthier, along with those with a diagnosed disease, could represent the greater proportion of workers attending health examinations because these were not compulsory. This causes bias in the recruitment procedure as, in addition, it is not well known whether the healthier workers or the ones with a diagnosed disease are the ones with the greatest interest in the checks. Nor can we ignore the bias of the healthy worker,
since those workers with serious illnesses would not be currently active. In addition, it is not well known if the healthiest workers or those with a diagnosed illness have the greatest interest in checks.

CONCLUSIONS
NIM-MetS has proved to be a valid method for the early detection of MetS in a healthy worker population. It is a simple, economical and quick non-invasive test which is easy to apply and interpret in any healthcare setting (primary healthcare, hospitals, occupational health) as well as in other settings (education, sport, etc). WHtR is the best predictor of MetS and its cut-off point can be used for both genders and for different age groups. The clinical decision tree that produces NIM-MetS uses WHtR (0.56) and BP (128/80 mm Hg), and obtains high SP and diagnostic validity. NIM-MetS provides a gradient or risk scale which allows a more accurate and earlier detection of CVD in subjects with risk of MetS.

Author affiliations
1Department of Occupational Safety and Health, Córdoba City Hall, Córdoba, Spain
2Department of Fundamental Biology and Health Sciences, University of the Balearic Islands, Palma, Spain
3Research Group on Evidence, Lifestyles & Health, Research Institute on Health Sciences (IUNICS), University of the Balearic Islands, Palma, Spain
4Department of Occupational Risk Prevention and Environmental Protection, University of Córdoba, Córdoba, Spain
5Prevention of Occupational Risks in Health Services, Balearic Islands Health Service, Palma, Spain
6IMIBIC, Reina Sofia University Hospital, University of Córdoba, Córdoba, Spain
7CIBER Physiopathology of Obesity and Nutrition CIBEROBN, ISCIII, Madrid, Spain
8Nursing and Physiotherapy Department, University of the Balearic Islands, Palma, Spain
9Department of Nursing, School of Medicine and Nursing, University of Córdoba, Córdoba, Spain

Contributors MR-S and contributed the conception, design, acquisition and analysis/interpretation of data, and drafted the manuscript. PT contributed to the data collection and analysis. MV-Contributed to analysis and interpretation of data. AA-AL-G contributed to the data collection and analysis. F-JF-J contributed to the analysis of data. AA and CA-F contributed to the conception. GM-R contributed to the analysis and interpretation of data. All authors critically revised the manuscript, gave final approval and agree to be accountable for the integrity and accuracy of all aspects of the work.

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