Introduction

The capacity of an ecosystem to tolerate disturbances without changing its original structure, functions and processes has been defined as ecological resilience [1,2]. Ecological systems can undergo regime shifts when disturbance thresholds are exceeded and reorganize in alternative states with new structures, functions and processes [3,4]. Ecological consequences of regime shifts are uncertain, and sometimes they are considered to have negative consequences for biodiversity and ecosystem service provisioning to humans; for instance, when cultural eutrophication triggers a shift from a clear-water to a turbid-water state [4,5]. After regime shifts, the new or alternative states can be stable, meaning that they resist returning to a state that existed prior to the regime shift [6]. In such cases, costly management and restoration interventions are needed to return ecosystems to resilient desired states [7,8,9,10].

Clear examples of ecosystems that are trapped in a degraded state are acidified lakes [11,12,13,14]. Although international agreements have resulted in reduced sulfur emissions to reduce acidification of freshwaters, several factors contribute to maintain lakes in an acidified state. Chemical weathering processes in soils [15,16], biological interactions [17], biotic resistance [18], food web stability [19], limited dispersal and population connectivity [20], scale-specific processes [21], and Allee effects [12,22] have been shown to constrain recovery to a desired, pre-acidification lake state. Thus, to protect sensitive biodiversity elements, especially the fish fauna, extensive liming programs have been established in many regions, including Sweden [23,24,25,26,27,28]. Despite the consequences of liming on aquatic biota and ecosystem properties being increasingly understood [29,30], it is unknown how liming affects the resilience of aquatic ecosystems.

Here, we describe an analysis of the resilience of phytoplankton community structure (resilience of what, [31]) in circum-neutral, acidified and limed lakes (i.e., lakes with different anthropogenic stress and management histories; resilience to what). We compare resilience characteristics of limed lakes relative to the degraded, undesired lake conditions (acidified lakes) and targeted reference conditions (circum-neutral lakes) [e.g., [12]], and more generally the efficiency of liming as a management tool. For such analyses, phytoplankton communities are appealing compared to other...
Within and cross-scale structures in ecosystems are related to diversity [46]. To evaluate how diversity characteristics influence the within- and cross-scale structure of acidified, limed and circum-neutral lakes, we first evaluate metrics of community structure that are commonly used in ecology, followed by time series modeling. Based on our current ecological knowledge of limed and acidified lakes, we test the hypothesis that the relative resilience of limed lakes is lower relative to circum-neutral and acidified lakes. This lower resilience should manifest in a reduced within- and cross-scale structure arising from lime applications that we expect shall disrupt and homogenize natural community assembly processes. Results may provide a mechanistic basis for understanding why communities return to an acidified state once liming is discontinued [30].

Materials and Methods

Ethics Statement

All field sampling and laboratory analyses reported in this study are part of either the Swedish National Lake Monitoring Program or the national monitoring program for Integrated Studies of the Effects of Liming of Acidified Waters, both regulated by the Swedish Agency for Marine and Water Management. All data are made freely available to the public via the web by the data host (Department of Aquatic Sciences and Assessment; Swedish University of Agricultural Sciences; www.slu.se/aquatic-sciences) and no permission for use of the data is therefore required. It is also confirmed that the field studies did not involve endangered or protected species.
Lake Selection

For this study, nine lakes representing three different management types were selected from the webpage hosting data from monitoring programs of Swedish inland waters (http://miljodata.slu.se/mvm/): 1) managed lakes (Ejgdesjön, Gyslattajon, Gysltigesjon) that were limed to mitigate the effects of anthropogenic acidification; 2) acidified lakes (Brunnsjön, Härsvattnet, Roteshogstajnen) that were unmanaged; 3) circum-neutral lakes (Allgjuttern, Fracksjon, Stora Skarsjon), used as reference lakes with “desired” ecosystem properties. The study lakes are situated in the same ecoregion (i.e. the boreonemoral ecoregion of southern Sweden), are of similar size (mean lake surface area of 0.31 km², range 0.11–0.83 km²) and have been monitored regularly for surface water chemistry and phytoplankton for 13 years (1997–2009).

Figure 1. Comparison of the community metrics in the different lakes: (a) total biovolume, (b) species richness; (c) Shannon-Wiener diversity index; (d) evenness.

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Sampling
Phytoplankton and two water chemistry variables that are directly related to liming (pH and calcium concentration) were sampled monthly during the ice-free period at a mid-lake station in each lake. Water was collected at 0.5 m depth with a Plexiglas sampler and kept cool during transport to the laboratory for further analysis. All physicochemical analyses were done by SWEDAC certified laboratories (Swedish Board for Accreditation and Conformity Assessment, SWEDAC; http://www.swedac.se/en/) at the Department of Aquatic Sciences and Assessment, Swedish University of Agricultural Sciences, following International (ISO) or European (EU) standards when available [47]. A broader characterization of the abiotic environment of the lakes studied is shown in Electronic Table S1. Standard sampling protocols for abiotic and biological variables were used throughout the study period.

Phytoplankton was sampled by taking a water sample from the epilimnion using a 2-m long Plexiglas tube sampler (diameter = 3 cm). In lakes with a surface area >1 km² a single mid-lake site was used for sampling. In lakes with a surface area <1 km², five random epilimnetic water samples were taken and mixed to form a composite sample from which a subsample was taken and preserved with acid Lugol's iodine solution [48]. Phytoplankton counts were made using an inverted light microscope and the modified Utermöhl technique commonly used in the Nordic countries [48]. Taxa were identified to the lowest taxonomic unit possible (usually species). Biovolumes (mm³ L⁻¹) were calculated from geometric shapes following protocols developed by Blomqvist and Herlitz [49].

Statistical Analyses
Community and Water Quality Analysis
Prior to analyses, phytoplankton data were averaged to obtain three values that were spaced nearly equidistantly in time, covering early spring, summer and late autumn each year. We characterized phytoplankton community structure across lakes using common metrics (total biovolume, richness, diversity and evenness) following recent recommendations by Jost [50] and Tuomisto [51,52] to obtain mathematically and statistically unbiased measures. The exponentiated Shannon index [50], which considers both species richness and evenness was used as a measure of “diversity” [51]. Exponentiation of the Shannon index expresses diversity in terms of species equivalents, making “diversity” and “richness” patterns directly comparable [50,51]. Evenness was obtained by dividing “diversity” with “richness” and therefore unrelated to richness [50,52].

Repeated measures analysis of variance (rm-ANOVA) was carried out in Statistica v.5 (Statsoft Inc, Tulsa, OK, USA) to test for significant differences in phytoplankton community metrics between the three categories of lakes. Similar rm-ANOVAs were conducted for pH and calcium concentration, to associate phytoplankton community dynamics with abiotic effects of liming. We tested for the effects of “management type” (circum-neutral, acidified and limed) (fixed factor), “time” (random factor) and their interactions. These factors comprised the independent variables while the community metrics comprised the dependent variables in the analysis. All data were log-transformed when necessary prior to analyses to fulfill the requirements of parametric tests. Because assumptions of sphericity were violated, degrees of freedom were corrected using the Greenhouse-Geisser adjustment.

### Table 2. Results of PERMANOVA contrasting phytoplankton communities between lakes (circum-neutral, acidified and limed), time (year) and their interactions.

| Metrics       | Statistics | Treatment | Time   | Treatment × Time |
|---------------|------------|-----------|--------|------------------|
| Communities   | d.f: 2, 6  | 38, 228   | 76, 228|                  |
|               | MS: 8.28   | 0.27      | 0.14   |                  |
|               | F: 38.81   | 1.24      | 0.65   |                  |
|               | P: <0.001  | <0.001    | 1      |                  |

Significant terms are emphasized in bold. (df: degrees of freedom, MS: mean squares, F ratio and P levels).

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### Table 3. Results of repeated-measures ANOVA contrasting water quality variables (pH and calcium concentration) between lakes (circum-neutral, acidified and limed), time (year) and their interactions.

| Parameters | Statistics | Treatment | Time   | Treatment × Time |
|------------|------------|-----------|--------|------------------|
| pH         | d.f: 2, 6  | 38, 228   | 76, 228|                  |
|            | MS: 80.25  | 0.05      | 0.03   |                  |
|            | F: 894.9   | 0.56      | 0.31   |                  |
|            | P: <0.001  | 0.98      | 1.00   |                  |
| Calcium (meq/L) | d.f: 2, 6  | 38, 228   | 76, 228|                  |
|            | MS: 1.30   | <0.01     | <0.01  |                  |
|            | F: 463.07  | 0.78      | 0.42   |                  |
|            | P: <0.001  | 0.83      | 1.00   |                  |

Significant terms are emphasized in bold. (df: degrees of freedom, MS: mean squares, F ratio and P levels).

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freedom were corrected following the procedure by Huynh and Feldt [53] (note that this adjustment can lead to degrees of freedom with decimals). Inference was made at $P < 0.05$. Tukey’s HSD test was performed to make a posteriori, group-wise comparisons of lake management types when a significant treatment effect was observed. We consider significant interaction terms between management type and time crucial for inferring differences in phytoplankton community metrics.

These univariate comparisons were complemented with multivariate analyses on phytoplankton communities using permutational multivariate analysis of variance in PERMANOVA version 1.6 [54]. PERMANOVA was based on a similar design as the univariate ANOVAs, using square root transformed species biovolume matrices that were converted in Bray–Curtis dissimilarity matrices and 9999 unrestricted permutations of raw data. Significant differences were inferred at an $\alpha$-level of 0.05.

Figure 2. Selected water quality variables: (a) pH and (b) calcium concentration between 1997 and 2009 in circum-neutral acidified and limed lakes. Shown are the overall patterns (mean ± SE) for the different lakes.
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Multivariate Time Series Modeling

We assessed two attributes of relative resilience (cross-scale and within-scale structure) of the phytoplankton communities with a time series modeling approach based on redundancy analysis (RDA) [42]. For an outline of the approach see the flow chart in Appendix S1 and Angeler et al [45]. We used temporal variables extracted by Asymmetric Eigenvector Maps (AEM) analysis [55,56]. Briefly, the AEM analysis produces a set of orthogonal temporal variables that are derived from the linear time vector that comprises the length of the study period (i.e., 39 time steps for each lake) and that can be used as explanatory variables to model temporal relationships in community data. The type of AEM variables computed in the present study was designed for spatial analysis to account for linear trends in the response variables. As time comprises a directional process, AEM is better suited to model linear trends relative to other methods (Principal Coordinates of Neighbor Matrices and Moran Eigenvector Maps [56]).
This procedure yielded AEM variables with positive Eigenvalues, each of which corresponds to a specific temporal structure and scale; the first AEM variable models linear trends and the subsequent variables capture temporal variability from slow to increasingly shorter fluctuation frequencies in the community data [57]. For each lake we constructed a parsimonious temporal model by running a forward selection on the AEM variables. Because AEM analysis is efficient in covering linear trends no detrending of models was necessary.

Redundancy analysis (RDA) retains significant AEM variables and these are linearly combined in ways to extract temporal patterns from the Hellinger-transformed species matrices (this transformation is achieved by dividing the species biovolumes by the row sum and taking the square root of the resulting values). The RDA identifies species with similar temporal patterns in the species × time matrix and uses their temporal pattern to calculate a modeled species group trend for these species based on linearly combined AEMs. The significance of the temporal patterns of modeled fluctuation patterns of species groups revealed by the RDA is tested by means of permutation tests. The RDA relates each modeled temporal fluctuation pattern with a significant canonical axis. The R software generates linear combination (lc) score plots, which visually present the modeled temporal patterns of species groups that are associated with each canonical axis. Based on the number of significant canonical axes, the number of modeled fluctuation patterns of species groups with independent temporal patterns can be deduced. The ecological relevance of each temporal pattern identified can be quantified, using adjusted $R^2$ values of the canonical axes. The overall temporal structure of the whole community can then be deduced from the number of significant canonical axes in the RDA models. The number of canonical (RDA) axes identified gives insight about the number of temporal scales at which phytoplankton community fluctuations take place and can therefore be used to assess the cross-scale structure attribute of resilience [44,45].

All relevant steps in the analyses were carried out with two functions implemented in the R 2.15.2 statistical software package [58]. First, the conversion of the linear time vector to AEM variables is done using the “aem.time” function (AEM package). All remaining steps (calculation of modeled species group trends, visual presentation of the results in form of lc score plots) are carried out with the “quickPCNM” function (PCNM package). All models are calculated exclusively based on an automatic statistical procedure, which limits bias in modeling scales that can be introduced by researcher subjectivity.

After identifying the cross-scale structure in phytoplankton community dynamics, we evaluated the within-scale attribute of resilience using correlation analysis. Spearman’s rank correlation analysis was used to investigate the relationship between individual phytoplankton species (raw biovolume data of individual species) with the linear combination (lc) scores extracted from significant canonical axes of the time series models for each lake type. With this approach we were able to calculate scale-specific taxon richness. Species that did not correlate with any canonical axes were considered to reflect stochastic dynamics. The number of these stochastic species was evaluated by subtracting the total number of species that correlated with canonical axes from the total number of species used for the time series modeling across lakes [45].

Results

Univariate and Multivariate Community Analyses

Univariate analyses of community metrics (Table 1; Figure 1) revealed significant lake management type effects for species richness and diversity but not for total biovolume and evenness. Post-hoc analysis revealed that the species richness and diversity were higher in circum-neutral lakes and limed lakes compared to acidified lakes (Tukey’s HSD test: circum-neutral = limed > acidified; P<0.05). The effect of time was significant for all metrics.
studied, but the interaction term (management type × time) was only significant for species richness. The results of the PERMANOVA identified a significant effect of management type and time whilst the interaction effect was not significant (Table 2). Group-wise analysis using PERMANOVA indicated that the phytoplankton community structure in the three different lake types were significantly different from each other (P<0.001).

**Water Chemistry: pH and Calcium Concentration**

The results of the repeated measures analysis of variance revealed that management type has an effect on both the pH and the calcium concentration, while the effect of time and the interaction term were not significant (Table 3 and Figure 2). Moreover, post-hoc analysis revealed that both the pH and the calcium concentration were higher in limed lakes than circum-neutral and acidified lakes, whereas circum-neutral lakes had higher values than acidified lakes (Tukey’s HSD test: Limed > circum-neutral > acidified; P<0.05).

**Multivariate Time Series Modeling**

Time series modeling revealed significant temporal structure in all lakes studied (Figure 3). The RDA models explained on average similar amounts of the adjusted variance of phytoplankton community dynamics (26% circum-neutral lakes, 27% acidified and 28% limed). The models show that the temporal fluctuation patterns of groups within the phytoplankton communities were unique in each lake independent of management practice. However, all circum-neutral lakes showed temporal dynamics at six significant temporal scales, compared to acidified (Brunnsmojn 6, Harsvattnet 4, Rotehogstjärnen 5) and limed lakes (Ejgdesjön 6, Gyldttesjön 5, Gyldttesjön 5) (Figure 3), highlighting a slightly higher and more consistent cross-scale structure in circum-neutral lakes relative to the other lake types. Using Spearman rank correlation analysis (Table 4) to assess the within-scale attribute of resilience, we found that the number of species contributing to the scale-specific temporal patterns was comparable across the different lake types, suggesting that they have similar within-scale structures. Species that did not correlate with any significant temporal frequency pattern (i.e. stochastic species) were on average higher in limed (27%) and acidified (23%) compared to circum-neutral (15%) lakes.

**Discussion**

The relative resilience of limed lakes in terms of the dynamic within- and cross-scale structure of phytoplankton communities was predicted to be lower relative to acidified and circum-neutral lakes. The within-scale component of resilience, expressed as the percentage of species that explained each temporal pattern, was similar across lake types. Also the cross-scale structure differed marginally between managed and unmanaged lakes, although water chemistry accounted for subtle differences in cross-scale structure. All circum-neutral lakes and the least acidified Brunnsmojn showed six distinct temporal frequency fluctuations, while the most acidified lake Härsvatten, that clearly comprises an undesired but stable state, had only four patterns. This pattern is counterintuitive because it suggests a lower resilience of this lake. This highlights that the cross-scale structure analysis, despite characterizing important features of complexity, may not capture the full spectrum of resilience. Our study shows how considering the dynamics of stochastic species can provide a broader picture of resilience. Similarly, despite liming substantially increasing pH and calcium concentration, no significant increase in cross-scale structure was observed, highlighting no pronounced effect of
management action on phytoplankton cross-scale attributes of resilience. Although our study hypothesis was rejected, our results provided us with a better understanding how management action, specifically liming, influences resilience.

Management and restoration efforts are often directed towards breaking equilibrium conditions of undesired system states, returning them to more desired states and increasing the resilience of these restored or managed states [7,8,9,10]. The finding of similar resilience characteristics across circum-neutral, acidified and limed lakes has several implications for management, especially regarding liming as a management tool [29,59,60].

The results can be interpreted in two contexts. (1) Circum-neutral lakes have been marginally affected by anthropogenic acidification because of their higher acid buffering capacity [61,62]. The resilience characteristics observed in those lakes therefore comprises the target of management. The similar within- and cross-scale features of resilience observed in the acidified lakes suggest that natural recovery has putatively led to approaching targeted (circum-neutral lakes) resilience conditions. In this case, given the similar within- and cross-scale resilience characteristics observed in limed lakes, we can conclude that costly management interventions do not necessarily achieve better conditions than unmanaged, natural recovery [60].

(2) The resilience characteristics observed in acidified lakes characterizes their basin of attraction and thus their resistance to return to desired target conditions [32]. This explanation seems more plausible, given that ecological recovery (i.e. resilience attributes and community composition) has not reached management targets in many cases despite decades of policy implementation [28,63]. This interpretation is also supported by research that has identified many abiotic and trophic factors that constraint recovery [12,13,64]. How is this related to the resilience observed in limed lakes?

The similar within- and cross-scale attributes of resilience in limed and acidified lakes, suggest that liming produces at least a partial management success in terms of increasing species richness and diversity, which has also been observed in other studies [29,65,66]. However, from a systemic point of view, it seems to reorganize communities within the acidified states rather than break the feedbacks that maintain lakes in the acidified state. This re-organization, which was manifested in communities that are neither representative of acidified nor circum-neutral lakes, was evident in our PERMANOVA analysis (see also NMDS plot in Appendix S2) and the water chemistry analysis (see also [29]).

From a resilience perspective, our results suggest that limed lakes comprise a particular condition of an acidified lake state; that is, liming keeps lakes in the “degraded” basin of attraction, thereby failing to restore and reinforce a desired state equivalent to pre-acidification conditions that maintains desired ecosystem attributes. Alternatively, if liming creates an alternative basin of attraction, it might be very shallow and therefore unstable. Whether or not limed lakes comprise a particular configuration of an acidified state or an alternative state itself, our results support findings that have suggested that communities return to an acidified state once liming is discontinued [30]. Thus, the broader ecological implications of liming lakes are the following: 1) when natural recovery of acidified lakes attains similar resilience characteristics as in targeted reference lakes, costly management practice might not be necessary, 2) ecologists have begun to regard liming as an ecosystem-level perturbation rather than an integral restoration tool [67,68,69]. In a more specific management context, liming may comprise some form of command and control management [70], whereby the partial success in terms of targeted increase in diversity and species richness, and habitat suitability to sustain fisheries might generate substantial negative side effects in the ecosystem due to biogeochemical alteration [71]. If liming only partly mitigates acidification impacts without fundamentally altering the equilibrium conditions of acidified states (i.e. restoring acidified lakes to pre-acidification conditions that become self-sustaining), further research will be required to foster our understanding of potentially negative side effects on the ecological integrity of managed lakes (e.g. altered Al toxicity [72], altered nutrient precipitation [73], material consumption by invertebrates [74], food web structure [29]).

In addition to assessing patterns of resilience across managed and unmanaged boreal lakes, our study provides new insight into the consequences of liming for biodiversity and its effects on resilience. Increased resilience has been associated with a higher species richness and diversity in communities [5,32]. While liming indeed increased the species richness and diversity in the limed lakes to targeted levels present in circum-neutral relative to acidified lakes, our time series models suggest that this increased diversity contributed little to within- and cross scale structure in the phytoplankton communities. This apparent paradox can be explained by the number of species with ostensibly stochastic dynamics that were on average higher in the limed and acidified compared to circum-neutral lakes. From a disturbance ecology perspective, our results are consistent with the findings of an increased importance of stochastic community assembly when ecosystems face perturbations [75,76]; if this pattern can be generalized, the argument that liming comprises a perturbation would find additional support. This also suggests that the most acidified lake Harsvatten with the lowest cross-scale structure detected, might be able to cope with disturbances, thereby maintaining its functions and feedbacks, because of the higher amount of stochastic species increasing adaptive capacity (e.g. a high response diversity [77]). Our study makes clear how the role of species richness can be scrutinized if partitioned into patterns that reflect both the deterministic and stochastic processes occurring at different scales [45], thereby highlighting the usefulness of time series modeling for assessing resilience. These patterns can be further explored for gaining a more process-based understanding of how management affects species diversity and their influence on resilience.

We conclude by highlighting that, given the similar within- and cross-scale attributes observed across lakes, an assessment of resilience characteristic of limed lakes would have been inconclusive. Without information of the broader ecological impacts of liming on communities and other ecosystem characteristics that has accumulated in the literature [29,30,68,69,71,72,73,74] and ecological knowledge of reference lakes, we would have not been able to judge whether lake liming achieved the ultimate management goal: of restoring and fostering desired ecosystem states. Assessments of the relative resilience of managed systems therefore require multiple lines of evidence, including reference sites that comprise management targets and approaches based on complex systems theory and “traditional” ways of characterizing community structure and functions to increase inference.

Supporting Information

Table S1 Summary of geographical positions, morphological characteristics and water chemistry of study lakes. Values represent the inter-annual mean value and standard deviation for the study period 1997–2009. (DOCX)
Appendix S1  Flow chart outlining the steps involved in time series modeling. (DOCX)

Appendix S2  Non-metric multidimensional scaling (NMDS) ordination showing phytoplankton communities across the different lake types over the study period (1997–2009). (DOCX)

References

1. Holling CS (1973) Resilience and stability of ecological systems. Annual Review of Ecology and Systematics 4: 1–23.

2. Bormann F, Likens GE (1979) Ecological self-organization. Annu. Rev. Ecol. Syst. 10: 109–140.

3. Beisner BE, Haydon DT, Cuddington K (2003) Alternative stable states in ecology. Frontiers in Ecology and the Environment 1: 4–14.

4. Scheffer M, Carpenter S, Foley JA, Folke C, Hannon J (2001) Catastrophic shifts in ecosystems. Nature 413: 591–596.

5. Folke C, Carpenter S, Walker B, Scheffer M, Elmqvist T, et al. (2004) Regime shifts, resilience, and biodiversity in ecosystem management. Annual Review of Ecology Evolution and Systematics 35: 557–581.

6. Scheffer M, Hosper SH, Meijer ML, Moss B, Jeppesen E (1993) Alternative equilibria in shallow lakes. Trends in Ecology & Evolution 8: 273–279.

7. Soltero R, Singleton L, Patmon C (1992) The changing sponge river watershed: actions to improve and maintain water quality. In: Naiman R, editor. Toward Sustainable Watershed Management. Springer New York. pp. 435–458.

8. Jeppeesen E, Merhoff DJ, Jacobsen BA, Hansen RS, Sondergaard M, et al. (2007) Restoration of shallow lakes by nutrient control and biomimicuration—the successful strategy varies with lake size and climate. Hydrobiologia 561: 269–285.

9. Cooke GD (1993) Restoration and management of lakes and reservoirs. Lewis Publishers (Boca Raton, Fla.).

10. Carpenter SR, Krafcik CE, Wright R, He X, Soranno PA, et al. (1992) Resilience and resistance of a lake phosphorus cycle before and after food web manipulation. The American Naturalist 140: 781–798.

11. Marmorek D, Korman J (1993) The use of zooplankton in a biomonitoring program to detect lake acidification and recovery. Water, Air, and Soil Pollution 69: 223–241.

12. Yan ND, Keller W, Somers KM, Pasono TW, Girard RE (1996) Recovery of crustacean zooplankton communities from acid and metal contamination: Comparing manipulated and reference lakes. Canadian Journal of Fisheries and Aquatic Sciences 53: 1301–1327.

13. Yan ND, Leung B, Keller W, Arnott SE, Gunn JM, et al. (2003) Developing conceptual frameworks for the recovery of aquatic biota from acidification. Ambio 32: 163–169.

14. Kopeck J, StuchlI M, Vesely J, Schaumburg J, Anderson I, et al. (2002) Acidification remediation: Evaluation of liming tests in small boreal lakes. Ambio 31: 36–44.

15. Bostedt G, Lofgren S, Innala S, Bishop K (2010) Acidification remediation in瑞典: Actions to improve and maintain water quality. In: Angeler DG, editor. Panarchy: Understanding Transformations in Human and Natural Systems. Washington D.C., USA: Island Press. pp. 376–382.

16. Jacks G, Joelsson A, Fleischer S (1994) Nitrogen retention in forest wetlands. Restor. Ecol. 2: 91–114.

17. Raitaniemi J, Rask M (1990) Preliminary observations on the effects of liming to acidified mountain lakes. Global Change Biology 16: 1189–1203.

18. Keller W, Yan ND (1998) Biological responses to liming in boreal lakes: an assessment using plankton, macroinvertebrate and fish communities. Journal of Applied Ecology 35: 478–486.

19. Layer K, Hildrew AG, Monteith D, Woodward G (2010) Long-term variation in zooplankton diversity across contrasting boreal lakes. Freshwater Biology 55: 1241–1256.

20. Kopaˇcˇek J, StuchlI ´k E, Vesely ´ J, Schaumburg J, Anderson I, et al. (2002) Acidification remediation: Evaluation of liming tests in small boreal lakes. Ambio 31: 36–44.

21.层器 J, 雪山 H, 竹林 R, 安德森 I, et al. (2002) Hysteretic Reversal of Central European Mountain Lakes from Atmospheric Acidification. Water, Air and Soil Pollution 90: 329–341.

22. Van ND, Keller W, Somers KM, Pasono TW, Girard RE (1996) Recovery of crustacean zooplankton communities from acid and metal contamination: Comparing manipulated and reference lakes. Canadian Journal of Fisheries and Aquatic Sciences 53: 1301–1327.

23. Van ND, Leung B, Keller W, Arnott SE, Gunn JM, et al. (2003) Developing conceptual frameworks for the recovery of aquatic biota from acidification. Ambio 32: 163–169.

24. Kopeck J, StuchlI M, Vesely J, Schaumburg J, Anderson I, et al. (2002) Acidification remediation: Evaluation of liming tests in small boreal lakes. Ambio 31: 36–44.
50. Jost L (2007) Partitioning diversity into independent alpha and beta components. Ecology 88: 2427–2439.
51. Tuomisto H (2010) A diversity of beta diversities: straightening up a concept gone awry. Part 1. Defining beta diversity as a function of alpha and gamma diversity. Ecography 33: 2–22.
52. Tuomisto H (2012) An updated consumer’s guide to evenness and related indices. Oikos 121: 1203–1218.
53. Huynt H, Feldt LS (1976) Estimation of the box correction for degrees of freedom from sample data in randomized block and split-plot designs. Journal of Educational Statistics 1: 69–82.
54. Anderson MJ (2001) A new method for non-parametric multivariate analysis of variance. Austral Ecology 26: 32–46.
55. Blanchet FG, Legendre P, Borcard D (2008) Modelling directional spatial processes in ecological data. Ecological Modelling 215: 325–336.
56. Blanchet FG, Legendre P, Maranger R, Monti D, Pepin P (2011) Modelling the effect of directional spatial ecological processes at different scales. Oecologia 166: 357–368.
57. Carpenter SR, Cole JJ, Pace ML, Batt R, Brock WA, et al. (2011) Early Warnings of Regime Shifts: A Whole-Ecosystem Experiment. Science 332: 1079–1082.
58. RDevelopmentCoreTeam (2012) R: A Language and Environment for Statistical Computing. Vienna, Austria: R Foundation for Statistical Computing.
59. Bradley DC, Ormerod SJ (2002) Long-term effects of catchment liming on invertebrates in upland streams. Freshwater Biology 47: 161–171.
60. Ormerod SJ, Durance I (2009) Restoration and recovery from acidification in upland Welsh streams over 25 years. Journal of Applied Ecology 46: 164–174.
61. Scott KA, Wielgolaski F, Gibbons J, Birks SJ (2010) Chemical characteristics and acid sensitivity of boreal headwater lakes in northwest Saskatchewan. Journal of Limnology 69: 33–44.
62. Lewis WM (1982) Changes in pH and buffering capacity of lakes in Colorado Rockies. Limnology and Oceanography 27: 167–172.
63. Henrikson L, Brodin Y-W (1995) Liming of surface waters in Sweden—a synthesis. Liming of Acidified Surface Waters: Springer. pp. 1–44.
64. Bink JA, Arnott SE, Sprules WG (2005) Local factors and colonist dispersal influence crustacean zooplankton recovery from cultural acidification. Ecological Applications 15: 2025–2036.
65. Jarvenc M, Kogppannski K, Rank M (1995) Responses of phyto- and zooplankton to liming in a small acidified humic lake. Water, Air, and Soil Pollution 85: 943–948.
66. Eriksson F, Hornstrom E, Mossberg P, Nyberg P (1985) Ecological effects of lime treatment of acidified lakes and rivers in Sweden. In: Forberg C, Johansson JA, editors. Forest Water Ecosystems: Springer Netherlands. pp. 145–163.
67. Bishop K, Laidon H, Hruska J, Kramp M, Kohler S, et al. (2001) Does acidification policy follow research in Northern Sweden? The case of natural acidity during the 1990’s. Water, Air, and Soil Pollution 130: 1415–1420.
68. McKie BG, Petrin Z, Malnquist B (2006) Mitigation or disturbance? Effects of liming on macroinvertebrate assemblage structure and leaf-litter decomposition in the humic streams of northern Sweden. Journal of Applied Ecology 43: 780–791.
69. Weatherley NS (1988) Liming to mitigate acidification in freshwater ecosystems: A review of the biological consequences. Water, Air, and Soil Pollution 39: 421–457.
70. Holling CS, Meffe GK (1996) Command and control and the pathology of natural resource management. Conservation Biology 10: 328–337.
71. Schindler DW (1997) Liming to restore acidified lakes and streams: A typical approach to restoring damaged ecosystems? Restoration Ecology 5: 1–6.
72. Roseland BO, Blakar IA, Bulger A, Kroghd F, Kvellestad A, et al. (1992) The mixing zone between limed and acidic river waters - complex Aluminum chemistry and extreme toxicity for salmonids. Environmental Pollution 78: 3–8.
73. Wallisdeit T, Berg H, Mehl M, Morth CM (2008) Influence of liming on metal sequestration in lake sediments over recent decades. Science of the Total Environment 407: 405–417.
74. Kullberg A, Petersen RC (1987) Dissolved organic carbon, seston and macroinvertebrate drift in an acidified and limed humic stream. Freshwater Biology 17: 535–564.
75. Kreyling J, Jeutsh A, Beierkuhnlein C (2011) Stochastic trajectories of succession initiated by extreme climatic events. Ecology Letters 14: 758–764.
76. Newbery DM, Lingenfelder M (2009) Plurality of tree species responses to drought perturbation in Bornean tropical rain forest. Plant Ecology 201: 147–167.
77. Mori AS, Furukawa T, Sasaki T (2013) Response diversity determines the resilience of ecosystems to environmental change. Biological Reviews 88: 349–364.