Abstract

We consider the impulse control of Lévy processes under the infinite horizon, discounted cost criterion. Our motivating example is the cash management problem in which a controller is charged a fixed plus proportional cost for adding to or withdrawing from his/her reserve, plus an opportunity cost for keeping any cash on hand. Our main result is to provide a verification theorem for the optimality of control band policies in this scenario. We also analyze the transient and steady-state behavior of the controlled process under control band policies and explicitly solve for an optimal policy in the case in which the Lévy process to be controlled is the sum of a Brownian motion with drift and a compound Poisson process with exponentially distributed jump sizes.

Keywords: impulse control; cash management problem; Lévy processes; Brownian motion.

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1. Introduction

Impulse control problems have a long history related to applications to the cash management problem. In the present paper, we consider the impulse control of Lévy processes. Our motivating application is the cash management problem in which there exists a system manager who must control the amount of cash he/she has on hand. We assume that the manager’s cash on hand fluctuates due to randomly occurring withdrawals from and deposits to his/her account but that the manager is charged a fixed plus proportional cost for any specific, intentional adding to or withdrawing from his/her reserves and that there exists an opportunity cost for keeping too little or too much cash on hand. The manager’s objective is to minimize his/her long run opportunity cost of keeping cash on hand plus any cost incurred from depositing or withdrawing from the reserve. An alternative motivating application which is also considered in the literature is a manager who wishes to control his/her inventory level. The manager’s inventory level fluctuates randomly and he/she may increase or decrease his/her inventory level at will by expediting or salvaging parts, paying a fixed plus proportional cost to do so. The manager’s objective is to minimize his/her long run inventory holding costs plus costs of expediting and salvaging.

Our first main result in the paper is to provide a verification theorem for the optimality of control band policies for the impulse control of Lévy processes. Our result is fairly general and holds for a wide class of opportunity cost functions. We then explicitly calculate the Laplace transform with respect to time and steady-state distribution of any spectrally positive Lévy process controlled under a control band policy. In Section 5, we consider the special case of a Lévy process which is comprised of the sum of a Brownian motion and a compound Poisson process with exponentially distributed jump sizes. In this specific case, we
show how one may use the results derived in this paper in order to characterize the value function for the associated control problem, and characterize the band levels in an optimal control policy as a solution of a system of equations. Moreover, we also show how one may determine the steady-state distribution of the controlled process when the underlying Lévy process is the sum of a negative drift and a compound Poisson process with exponentially distributed jump sizes.

The technique of impulse control was originally developed by Bensoussan and Lions [3, 4] and extended by Richard et al. [8, 19] Harrison, Selke and Taylor [11] and Sulem [21] also consider the impulse control of Brownian motion and explicitly calculate the critical parameters determining the optimal policy. In [10], an iterative computational scheme is provided in order to determine an optimal policy for the impulse control of Brownian motion. Recently, Ormeci, Dai and Vande Vate [17] have considered impulse control of Brownian motion under the average cost criterion and again show that a control band policy is optimal. Cadenillas, Zapatero, and Sarkar [7] and Cadenillas, Lakner, and Pinedo [6] solved the Brownian case with a mean-reverting drift. None of the above mentioned works allow there to be jumps in the process to be controlled. In [5], the optimality of an $(s, S)$ policy is proven for a process which is the sum of a constant drift, a Brownian motion, and a compound Poisson process. In related work, Bar-Ilan, Perry and Stadje [2] have also considered the problem of impulse control of Lévy processes for the specific case in which the Lévy process is a sum of a Brownian motion and a compound Poisson process. Assuming that a control band policy is optimal, their main results are to evaluate the cost functionals of the resulting policy through a fundamental identity derived from the martingale originally introduced by Kella and Whitt [14]. In [23], Yamazaki uses a scale function approach to study one-sided impulse control problems for spectrally positive Lévy demand processes. In [9], an impulse control problem is studied for a refracted Lévy process where the ruin time is modeled by a Parisian delay. Finally, [22] studies the impulse control of a geometric Lévy process.

2. The Model

In this Section we provide the specifics of the model described in the Introduction. All forthcoming processes are assumed to live on a probability triplet equipped with a filtration $\mathcal{F} = \{\mathcal{F}_t, 0 \leq t < \infty\}$. We begin by assuming that $Y_t$ is a Lévy process started from $x$ with Lévy measure $\nu$ such that

$$\int_{\{|y|\geq 1\}} |y| \nu(dy) < \infty. \quad (1)$$

The process $Y_t$ will be used to represent the cash on hand process assuming that the manager exerts no control by making no deposits to or withdrawal from his/her fund. Let $J(\omega, dt, dy) = J(dt, dy)$ be the jump measure of $Y$. Then, $Y_t$ has the Itô-Lévy decomposition

$$Y_t = x + \mu t + \sigma w_t + A_t + \varphi_t \quad (2)$$

where $\varphi_t$ is the following martingale

$$\varphi_t = \int_{\{|x|<1\}} x \{J((0, t], dx) - t \nu(dx)\}$$

and $A$ is the sum of the “large” jumps

$$A_t = \sum_{0<s\leq t} \Delta Y_s 1_{\{\Delta Y_s \geq 1\}}.$$ 

The process $w$ is assumed to be a standard Wiener process and $\mu$ is a constant. We do not make any assumption regarding $\sigma$, it may be zero or non-zero. Also we allow $\nu(\mathbb{R}) = 0$ in which case $Y$ is continuous. The case when both $\sigma = 0$ and $\nu(\mathbb{R}) = 0$ is also included, although this case is trivial ($Y$ is deterministic in this case). In general, we will use $P_x$ to denote the probability measure under which $Y_t$ is started from $x$. 


and $E_x$ its associated expectation operator. We note that $\varphi$ is a quadratic pure jump local martingale ([18], Chapter II, Section 6) with quadratic variation

$$[\varphi]_t = \sum_{0 \leq s \leq t} (\Delta Y_s)^2 1_{(|\Delta Y_s| \leq 1)} = \int_{(0,t] \times \mathbb{R}} y^2 1_{(|y| < 1)} J(ds, dy),$$

which has expected value

$$E[\varphi]_t = t \int_{\mathbb{R}} y^2 1_{(|y| < 1)} \nu(dy) < \infty. \quad (3)$$

It follows ([18], Chapter II, Corollary 3 to Theorem 27) that $\varphi$ is a square-integrable martingale.

We let

$$T, \Xi = (\tau_1, \tau_2, \ldots, \tau_n, \ldots, \xi_1, \xi_2, \ldots, \xi_n, \ldots)$$

denote the impulse control policy used by the manager where $0 \leq \tau_1 < \tau_2 < \tau_3 \ldots$ are $[0, \infty]$-valued stopping times and $\xi_n$ is an $F_{\tau_n}$ measurable random variable for each $n \geq 1$. Positive values of $\xi_n$ represent deposits by the manager into his/her fund and negative values represent withdrawal. We allow the possibility that $P_x(\tau_n = \infty) > 0$. In that case on the event $\{\tau_n = \infty\}$ there are fewer than $n$ interventions on the time horizon $[0, \infty)$. We require that $\xi_n = 0$ on $\{\tau_n = \infty\}$. However, we also require that $\xi_n \neq 0$ on $\{\tau_n < \infty\}$.

As described in the Introduction, the controlled cash on hand process $X_t$ follows the dynamics

$$X_t = Y_t + \sum_{i=1}^{\infty} 1_{\{\tau_i \leq t\}} \xi_i \quad (4)$$

and has RCLL paths. Let $\lambda > 0$ be a fixed discount factor.

**Definition 2.1.** A control policy $(T, \Xi)$ is called admissible if for some constant $K \in (0, \infty)$ the following relation holds:

$$e^{-\lambda t}|X_t| < K, \quad \text{a.s., } t \in [0, \infty). \quad (5)$$

We emphasize that the constant $K$ is not “universal”; for different policies it may be different.

Let the opportunity cost function be a right-continuous function $\phi : \mathbb{R} \mapsto [0, \infty)$, so the cost for cash on hand being at level $x$ is $\phi(x)$. We shall assume that for some constants $K_1, K_2 > 0$

$$K_1 \phi(x) + K_2 \geq |x|, \quad x \in \mathbb{R}. \quad (6)$$

The manager’s total cost is the sum of his/her expected discounted opportunity costs as well as impulse control costs and is given by

$$I(x, T, \Xi) = E_x \left[ \int_0^\infty e^{-\lambda t} \phi(X_t) dt + \sum_{n=1}^{\infty} e^{-\lambda \tau_n} g(\xi_n) \right] \quad (7)$$

where the manager’s impulse control costs are given by

$$g(\xi) = \begin{cases} 
C + c\xi, & \text{if } \xi > 0, \\
0, & \text{if } \xi = 0, \\
D - d\xi, & \text{if } \xi < 0. 
\end{cases} \quad (8)$$

We assume that the fixed costs $C$ and $D$ are positive and the variable costs $c$ and $d$ are non-negative constants. The term $e^{-\lambda \tau_n}$ in (7) is well defined once we set $e^{-\infty} = 0$.

One of our primary objectives in this paper is to identify an optimal impulse control $(T, \Xi)$ that minimizes the above total cost $I(x, T, \Xi)$. The value function of this optimization problem is

$$V(x) = \inf \{I(x, T, \Xi), \ (T, \Xi) \text{ is an admissible impulse control}\}.$$
Lemma 2.2. Let $(T, \Xi)$ be a control such that $I(x, T, \Xi) < \infty$. Then
\[
\lim_{n \to \infty} \tau_n(\omega) = \infty, \quad \text{a.s.}
\]
and
\[
\liminf_{t \to \infty} E_x \left[ e^{-\lambda t} |X_t| \right] = 0
\]

Proof: We prove first (9). Suppose the opposite, i.e., that $P(G) > 0$ where $G = \{ \omega \in \Omega : \tau_n \uparrow \tau < \infty \}$. Then
\[
\sum_{n=1}^{\infty} E_x \left[ e^{-\lambda \tau_n} g(\xi_n) \right] \geq \sum_{n=1}^{\infty} E_x \left[ e^{-\lambda \tau_n} g(\xi_n) 1_G \right] \geq \sum_{n=1}^{\infty} E_x \left[ e^{-\lambda \tau_n} 1_G \right] \text{min} \{ C, D \}
\]
\[
\geq \sum_{n=1}^{\infty} E_x \left[ e^{-\lambda t} 1_G \right] \text{min} \{ C, D \} = \infty,
\]
which contradicts our assumption that $I(x, T, \Xi) < \infty$. The second inequality in the above chain holds because on the event $G$ we have $\tau_n < \infty$ for every $n$, hence $\xi_n \neq 0$. Next we prove (10). By (6)
\[
\int_0^\infty E_x \left[ e^{-\lambda t} |X_t| \right] dt \leq \int_0^\infty E_x \left[ e^{-\lambda t} (K_1 \phi(x_t) + K_2) \right] dt,
\]
which is finite by our assumption that $I(x, T, \Xi) < \infty$, and (10) follows. \hfill \bullet

In the section that follows we show that an optimal impulse control takes the form of a control band policy which arises frequently as the solution to impulse control problems. Moreover, in Section 5 we provide an example in which we are able to explicitly identify the parameters corresponding to this policy.

3. Main Results

In this Section, we provide the main results of the paper, Theorems 3.1 and 3.2, showing that a solution of an ordinary differential equation that satisfies some additional conditions, must be the value function $V$. Also included in the statement of Theorem 3.2 is an optimal impulse control policy which turns out to be a double bandwidth control policy. We begin first with some preliminary results before providing the statements of Theorems 3.1 and 3.2.

For a function $f : \mathbb{R} \mapsto \mathbb{R}$ we define the operator
\[
Mf(x) = \inf \{ f(x + \eta) + g(\eta), \ \eta \in \mathbb{R} \setminus \{0\} \}.
\]
We shall also use the linear operator $A$ associated with the uncontrolled process $Y$, that is, for $f \in C^2(\mathbb{R})$
\[
Af(x) = \frac{1}{2}\sigma^2 f''(x) + \mu f'(x) + \int_\mathbb{R} [f(x + y) - f(x) - f'(x)y 1_{\{|y| < 1\}}] \nu(dy).
\]
(11)

Our assumption (1) and Taylor’s theorem implies that the integral on the right-hand side is finite whenever $f'$ and $f''$ are bounded. Indeed,
\[
\left| \int_\mathbb{R} [f(x + y) - f(x) - f'(x)y 1_{\{|y| < 1\}}] \nu(dy) \right|
\leq \int_\mathbb{R} |f(x + y) - f(x)| 1_{\{|y| \geq 1\}} \nu(dy) + \int_\mathbb{R} |f(x + y) - f(x) - f'(x)y| 1_{\{|y| < 1\}} \nu(dy)
\leq \text{const} \times \left[ \int_\mathbb{R} |g| 1_{\{|y| \geq 1\}} \nu(dy) + \int_\mathbb{R} y^2 1_{\{|y| < 1\}} \nu(dy) \right]
< \infty.
\]
(12)

In order to prove our results we shall actually need to extend the domain of $A$ to a larger class of functions $\mathcal{D}$ defined below.
Definition 3.1. Let $\mathcal{D}$ be the class of functions $f : \mathbb{R} \mapsto \mathbb{R}$ for which there exist an integer $n \geq 0$ and a set of real numbers $S = \{x_1, x_2, \ldots, x_n\}$ (if $n = 0$ then $S$ is the empty set) such that the following conditions hold:

(i) $f \in C^1(\mathbb{R}) \cap C^2(\mathbb{R} \setminus S)$

(ii) The derivative $f'$ is bounded on $\mathbb{R}$ and the second derivative $f''$ is bounded on $\mathbb{R} \setminus S$. We shall call the points in $S$ the exceptional points.

From Lemma Appendix A.2 follows that (12) holds even if $f \in \mathcal{D}$. Indeed, let $(f_n, n \geq 0)$ be a sequence of functions guaranteed by Lemma Appendix A.2, then we have

$$
\int_\mathbb{R} |f(x + y) - f(x) - f'(x)y| 1_{\{|y| < 1\}} \nu(dy)
= \int_\mathbb{R} \lim_{n \to \infty} |f_n(x + y) - f_n(x) - f'_n(x)y| 1_{\{|y| < 1\}} \nu(dy) \leq \text{const} \times \int_\mathbb{R} y^2 1_{\{|y| < 1\}} \nu(dy) < \infty.
$$

It follows that we can and will extend the operator $A$ to $\mathcal{D}$. This way $Af(x)$ may be undefined if $x$ is an exceptional point of $f$, but this will not cause any problems. For $f \in \mathcal{D}$ Itô’s rule applied to $f(X_t)$ holds in its usual form (see the Appendix A).

We now conjecture that an optimal impulse control policy takes a double bandwidth control policy form. In particular, we assume that there exist constants $a < \alpha \leq \beta < b$ such that

$$
\tau^*_1 = \inf \{t \geq 0 : Y_t \in \mathbb{R} \setminus (a, b)\}, \quad (13)
$$

and for $n \geq 2$

$$
\tau^*_n = \inf \{t \geq \tau^*_{n-1} : X_{t-} + \Delta X_t \in \mathbb{R} \setminus (a, b)\}. \quad (14)
$$

The jump sizes are given by the following equations. If $x \in \mathbb{R} \setminus (a, b)$ then

$$
\xi^*_1 = \begin{cases} 
\beta - x, & \text{if } x \geq b, \\
\alpha - x, & \text{if } x \leq a,
\end{cases} \quad (15)
$$

and for all other cases (including the case of $n = 1$ and $x \in (a, b)$)

$$
\xi^*_n = \begin{cases} 
\beta - (X(\tau^*_n) + \Delta Y(\tau^*_n)), & \text{if } X(\tau^*_n) + \Delta Y(\tau^*_n) \geq b, \\
\alpha - (X(\tau^*_n) + \Delta Y(\tau^*_n)), & \text{if } X(\tau^*_n) + \Delta Y(\tau^*_n) \leq a.
\end{cases} \quad (16)
$$

Note that $\tau^*_1 = 0$ if and only if $x \in \mathbb{R} \setminus (a, b)$. For $n \geq 1$ it is possible that $\Delta Y(\tau^*_n) = 0$ in which case $X(\tau^*_n)$ is either equal to $a$ or to $b$, $\xi^*_n$ is $\alpha - a$ or $\beta - b$, and $X(\tau^*_n)$ equal to $\alpha$ or $\beta$, respectively. However, it is also possible that $\Delta Y(\tau^*_n) \neq 0$, in which case $X(\tau^*_n) + \Delta Y(\tau^*_n)$ may be either larger than $a$ or smaller than $b$, but we still have $X(\tau^*_n)$ equal to $\alpha$ or $\beta$, respectively. This control policy is admissible since the corresponding cash at hand process is bounded, even without discounting. Such a policy corresponding the constants $a, \alpha, \beta, b$ will be denoted by $(T^*, \Xi^*)$.

Proposition 3.2. Suppose that for some $a < \alpha \leq \beta < b$ the sequence of stopping times $(\tau^*_n)_{n \geq 1}$ is given in (13)- (14). If $Y$ is not constant then $\tau^*_n < \infty$, $\tau^*_n < \tau^*_{n+1}$ for $n \geq 1$ almost surely, and $\lim_{n \to \infty} \tau^*_n = \infty$.

Proof: We are going to show that $\tau^*_n < \infty$, $P_2$-a.s. by induction. We have $\tau^*_1 < \infty$ because the sample paths of the Lévy process $Y$ are unbounded functions (Sato [20], Proposition 37.10). Suppose now that $\tau^*_{n-1} < \infty$ $P_2$-a.s. Notice that

$$
X_t - X(\tau^*_{n-1}) = Y_t - Y(\tau^*_{n-1}), \quad t \in [\tau^*_{n-1}, \tau^*_n), \quad (17)
$$

and on the event $\{\tau^*_n = \infty\}$ we have $|X_t - X(\tau^*_{n-1})| < b - a$ for $t \in [\tau^*_{n-1}, \infty)$, thus on $\{\tau^*_n = \infty\}$ every sample function $t \mapsto Y_t - Y(\tau^*_{n-1})$ is bounded by $b - a$ on $[\tau^*_{n-1}, \infty)$. By the strong Markov property for $Y$,
the process \( \{Y(\tau_n^* - 1) - s) - Y(\tau_n - 1), s \geq 0\} \) has the same law under \( P_x \) as \( \{Y_s - Y_0, s \geq 0\} \), and recalling again Proposition 37.10 in Sato [20] we conclude that \( P_x(\tau_n^* = \infty) = 0 \).

Next we show \( \tau_n^* < \tau_n^* + 1 \). Let \( Y^\alpha \) be the process \( Y \) started at \( Y_0 = \alpha \) and \( Y^\beta \) be the process \( Y \) started at \( Y_0 = \beta \). Let \( \tau_\alpha = \inf \{s \geq 0 : Y^\alpha \notin (a, b)\} \) and \( \tau_\beta = \inf \{s \geq 0 : Y^\beta \notin (a, b)\} \). By the right-continuity of \( Y \) we have \( \tau_\alpha > 0 \) and \( \tau_\beta > 0 \) a.s. By the strong Markov property for \( Y \), conditionally on \( X(\tau^*_n) = \alpha \) the inter-arrival time \( \tau_{n+1}^* - \tau_n^* \) has the same distribution as \( \tau_\alpha \) and conditionally on \( X(\tau^*_n) = \beta \) the inter-arrival time \( \tau_{n+1}^* - \tau_n^* \) has the same distribution as \( \tau_\beta \). Hence \( \tau_n^* < \tau_n^* + 1 \).

Finally we show that \( \lim n \to \infty \tau_n^* = \infty \). We take limits in (17) as \( t \to \tau_n^* \) and derive that

\[
X(\tau_n^* -) + \Delta Y(\tau_n^*) - X(\tau_n^* - 1) = Y(\tau_n^*) - Y(\tau_n^* - 1). \tag{18}
\]

Notice that the absolute value of the left-hand side is bounded below by \( \min(\alpha - a, b - \beta) \). Indeed, \( X(\tau_n^* -) + \Delta Y(\tau_n^*) \) is not included in the interval \((a, b)\), whereas \( X(\tau_n^* - 1) \) is either equal to \( \alpha \) or to \( \beta \). There are two cases. If \( X(\tau_n^* -) + \Delta Y(\tau_n^*) \leq \alpha \), then \( |X(\tau_n^* -) + \Delta Y(\tau_n^*) - X(\tau_n^* - 1)| \geq \alpha - a \). If \( X(\tau_n^* -) + \Delta Y(\tau_n^*) \geq b \), then \( |X(\tau_n^* -) + \Delta Y(\tau_n^*) - X(\tau_n^* - 1)| \geq b - \beta \). Then the absolute value of the right-hand side of (18) is also bounded below by \( \min(\alpha - a, b - \beta) \). Let \( \tau = \lim n \to \infty \tau_n^* \). It follows that the sample paths of \( Y \) do not have a limit on the left at \( \tau \) on \( \{\tau < \infty\} \), because from the property that \( Y(\tau_n^*) - Y(\tau_n^* - 1) \) is bounded below by a constant follows that \( Y(\tau_n^*), n \geq 1 \) is not a Cauchy sequence. This implies that \( P_x(\tau < \infty) = 0 \).

The following theorems contain the main results of this Section. The first theorem gives conditions under which a function is a lower bound for the value function of the optimization problem.

**Theorem 3.1.** Suppose that a function \( f : \mathbb{R} \to (0, \infty) \) such that \( f \in \mathcal{D} \) with a set of exceptional points \( S \) satisfies the following conditions:

(i) \( \mathcal{A} f(x) - \lambda f(x) + \phi(x) \geq 0 \) for \( x \in \mathbb{R} \setminus S \);

(ii) \( f(x) \leq M f(x) \) for \( x \in \mathbb{R} \).

Then \( f(x) \leq V(x) \).

The next theorem will characterize the value function and an optimal double band policy.

**Theorem 3.2.** Suppose that there exist constants \( a < \alpha \leq \beta < b \) and a function \( f : \mathbb{R} \to (0, \infty) \) such that \( f \in \mathcal{D} \) with a set of exceptional points \( \{a, b\} \), and the following conditions are satisfied:

(a) \( \mathcal{A} f(x) - \lambda f(x) + \phi(x) = 0 \) for \( x \in (a, b) \);

(b) \( f(x) \leq M f(x) \) for \( x \in (a, b) \);

(c) \( \mathcal{A} f(x) - \lambda f(x) + \phi(x) \geq 0 \) for \( x \in \mathbb{R} \setminus [a, b] \);

(d) \( f(a) = M f(a) = f(a) + C + c(\alpha - a) \), \( f(b) = M f(b) = f(\beta) + D + d(b - \beta) \);

(e) \( f \) is linear on \((-\infty, a]\) with slope \(-c\), and also linear on \([b, \infty)\) with slope \(d\).

Then \( f(x) = V(x) \). Furthermore, the control \((T^*, \Xi^*)\) given in (13), (14), (15), and (16) with these values for \( a, \alpha, \beta, b \) is optimal.

The requirements imposed on \( f(\cdot) \) in Theorem 3.2 are stronger than the ones imposed in Theorem 3.1, though it is not obvious since in Theorem 3.1 we required that \( f(x) \leq M f(x) \) on the entire of \( \mathbb{R} \), whereas in Theorem 3.2 we only required this inequality to hold on \([a, b]\) (see conditions (b) and (d)). However, it follows from the conditions of Theorem 3.2 that actually \( f(x) = M f(x) \) for \( x \in \mathbb{R} \setminus [a, b] \), as shown in the lemma below.

**Lemma 3.3.** Under the conditions of Theorem 3.2 we have that

\[
f(x) = M f(x) = f(\alpha) + C + c(\alpha - x), \quad x < a \quad \tag{19}
\]

and

\[
f(x) = M f(x) = f(\beta) + D + d(x - \beta), \quad x \geq b \quad \tag{20}
\]

**Proof:** We shall prove only (19); the proof of (20) is similar. First we note that we can calculate \( M f(a) \) by taking infimum over \((0, \infty)\) only, i.e.,

\[
M f(a) = \inf \{f(a + \eta) + g(\eta), \eta > 0\},
\]
because by condition (d) we have that $f(a) = Mf(a)$, and by condition (e) and by (8) we have that
\[ \inf\{f(a+\eta)+g(\eta), \eta < 0\} = f(a) + D > f(a). \]
Next let $x \leq a$. The function $\eta \mapsto f(x+\eta)+g(\eta)$ is constant on $(0, a-x]$. Indeed, if $\eta \in (0, a-x]$ then by condition (e) and by (8) we have that $f(x+\eta)+g(\eta) = f(x) + C+cx$. On the other hand $\inf\{f(x+\eta)+g(\eta), \eta < 0\} = f(x) + D$, hence using the notation $a \land b = \min\{a, b\}$ we have that
\[
Mf(x) = \inf\{f(x+\eta)+g(\eta), \eta \in \mathbb{R} \setminus \{0\}\} = (f(x) + D) \land \inf\{f(x+\eta)+g(\eta), \eta > a-x\}
\]
Since $f(x) + D > f(a)$, we have $Mf(x) = (f(x) + D) \land \inf\{f(x+\eta)+g(\eta), \eta > a-x\} = (f(x) + D) \land \inf\{f(x+\eta)+g(\eta), \eta > a-x\} = f(x) + D$. The last identity follows from condition (e), since under that condition $f(a) = f(a) + c(a-x)$.

Notice that by our conditions $\inf\{f(x+\eta)+g(\eta), \eta < 0\} = f(x) + D$, hence using the notation $a \land b = \min\{a, b\}$ we have that
\[
Mf(x) = \inf\{f(x+\eta)+g(\eta), \eta \in \mathbb{R} \setminus \{0\}\} = (f(x) + D) \land \inf\{f(x+\eta)+g(\eta), \eta > a-x\}
\]
Since $f(x) + D > f(a)$, we have $Mf(x) = (f(x) + D) \land \inf\{f(x+\eta)+g(\eta), \eta > a-x\} = (f(x) + D) \land \inf\{f(x+\eta)+g(\eta), \eta > a-x\} = f(x) + D$. The last identity follows from condition (e), since under that condition $f(a) = f(a) + c(a-x)$.

Proof of Theorem 3.1: Let $(T, \Xi) = \{\tau_1, \tau_2, \ldots, \xi_1, \xi_2, \ldots\}$ be an arbitrary admissible impulse control for which (9) and (10) hold. We are going to show that
\[
f(x) \leq I(x, T, \Xi).
\]
(21)
The assumption that (9) and (10) hold does not restrict the generality of this proof, since if any of (9) or (10) does not hold, then (21) is obvious by Lemma 2.2. By a similar calculation to (12) based on Taylor’s theorem one can show that
\[
\int_{(0,\ell] \times \mathbb{R}} e^{-\lambda s} \{f(X_{s-} + y) - f(X_{s-}) - f'(X_{s-})y1_{\{|y| < 1\}}\} J(ds, dy)
\]
\[
\leq \text{const} \times \left[ \int_{(0,\ell] \times \mathbb{R}} |y|1_{\{|y| \geq 1\}} J(ds, dy) + \int_{(0,\ell] \times \mathbb{R}} y^21_{\{|y| < 1\}} J(ds, dy) \right] < \infty.
\]
(22)
Notice that by our conditions $f \in D$, thus, by (12) and (22) the process $\{U_t, t \in [0, \infty)\}$ given by
\[
U_t = \int_{(0,\ell] \times \mathbb{R}} e^{-\lambda s} \{f(X_{s-} + y) - f(X_{s-}) - f'(X_{s-})y1_{\{|y| < 1\}}\} \{J(ds, dy) - ds \nu(dy)\}
\]
is well defined. This process is a local martingale by Corollary 11.10 and Theorem 11.45, part 3 in [12]. In order to apply this Corollary, we need that the process on the left-hand side of (22) has locally integrable variation; but it actually has integrable variation, since (22) holds even if we take the absolute value of the integrand on the left-hand side, and the expected value of the right-hand side is finite by (1) (see also [18], Theorems I.38 and I.44(ii)). Since $U$ is a local martingale, there exists a sequence of stopping times $(S_n)$ such that $\lim_{n \to \infty} S_n = \infty$, almost surely, and $\{U_{t \wedge S_n}, t \in [0, \infty)\}$ is a martingale for every $n$. Let $t \geq 0$ fixed, and let $T_n = \min\{S_n, t\}$. By the generalized Itô’s rule (Proposition Appendix A.1) and by the integration of parts formula (Protter, [18], Corollary 2 to Theorem II.22), we have that
\[
e^{-\lambda T_n} f(X_{T(n)}) - f(x) = \int_{(0, T_n]} e^{-\lambda s} f'(X_{s-}) dX_s + \int_0^{T_n} e^{-\lambda s} \left\{ \frac{a^2}{2} f''(X_s) - \lambda f(X_s) \right\} ds
\]
\[
+ \sum_{s \leq T_n} e^{-\lambda s} \{f(X_s) - f(X_{s-}) - f'(X_{s-})\Delta X_s\}.
\]
By (2) and (4) this can be cast in the form
\[
e^{-\lambda T_n} f(X_{T(n)}) - f(x) = \int_{(0, T_n]} e^{-\lambda s} f'(X_{s-})(dA_s + d\varphi_s + \sigma w_s)
\]
(23)
The sum in (28) can be written as the form

\[ \sum_{s \leq T_n} e^{-\lambda s} \{ f(X_s) \} \]

where \( \lambda \) is such that \( \tau_i - \lambda \).

The jumps of \( X \) consist of the jumps included in the control, thus we can cast (24) in the form

\[ \sum_{s \leq T_n} e^{-\lambda s} \{ f(X_s) - f(X_{s-}) - f'(X_{s-}) \Delta Y_s \} \]

which can be written as

\[ \sum_{s \leq T_n} e^{-\lambda s} \{ f(X_s) - f(X_{s-} + \Delta Y_s) \} + \sum_{s \leq T_n} e^{-\lambda s} \{ f(X_{s-} + \Delta Y_s) - f(X_{s-}) - f'(X_{s-}) \Delta Y_s \} . \]

Note that the first sum in the above expression has only finitely many terms; the number of terms is equal to the number of \( i \)'s such that \( \tau_i \leq T_n \), which is finite by (9). We substitute (25) into (24), and get

\[ e^{-\lambda T_n} f(X_{T(n)}) - f(x) = \int_{[0,T_n]} e^{-\lambda s} f'(X_{s-}) (dA_s + d\varphi_s + \sigma w_s) \]

or

\[ \int_{[0,T_n]} e^{-\lambda s} \left\{ \frac{\sigma^2}{2} f''(X_s) - \lambda f(X_s) + \mu f'(X_s) \right\} ds \]

\[ + \sum_{s \leq T_n} e^{-\lambda s} \{ f(X_s) - f(X_{s-}) - f'(X_{s-}) \Delta Y_s \} \]

\[ + \sum_{s \leq T_n} e^{-\lambda s} \{ f(X_{s-} + \Delta Y_s) - f(X_{s-}) - f'(X_{s-}) \Delta Y_s \} . \]

We cancel the integral with respect to \( A_s \) in (26) with the “large” jumps in (27), and write

\[ e^{-\lambda T_n} f(X_{T(n)}) - f(x) = \int_{[0,T_n]} e^{-\lambda s} f'(X_{s-}) (d\varphi_s + \sigma w_s) \]

\[ + \int_{[0,T_n]} e^{-\lambda s} \left\{ \frac{\sigma^2}{2} f''(X_s) - \lambda f(X_s) + \mu f'(X_s) \right\} ds \]

\[ + \sum_{s \leq T_n} e^{-\lambda s} \{ f(X_s) - f(X_{s-} + \Delta Y_s) \} \]

\[ + \sum_{s \leq T_n} e^{-\lambda s} \{ f(X_{s-} + \Delta Y_s) - f(X_{s-}) - f'(X_{s-}) \Delta Y_s \} . \]

The sum in (28) can be written as

\[ U_{T(n)} + \int_{[0,T_n] \times \mathbb{R}} e^{-\lambda s} \left\{ (f(X_{s-} + y) - f(X_{s-}) - f'(X_{s-}) y 1_{|y| < 1}) \right\} \nu(dy) ds, \]

so after substituting this back to (23)-(24), and using (11) we get the following equality:

\[ e^{-\lambda T_n} f(X_{T(n)}) - f(x) = \int_{[0,T_n]} e^{-\lambda s} f'(X_{s-}) (d\varphi_s + \sigma w_s) + \int_{[0,T_n]} e^{-\lambda s} \left\{ A f(X_{s-}) - \lambda f(X_{s-}) \right\} ds \]

\[ + \sum_{s \leq T_n} e^{-\lambda s} \{ f(X_s) - f(X_{s-} + \Delta Y_s) \} + U_{T(n)} . \]

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Next we take expectations; all martingale terms will disappear. Indeed, the local martingale
\[ Z_t = \int_{(0,t]} e^{-\lambda s} f'(X_s) d\phi_s \]
has quadratic variation
\[ [Z]_t = \int_{(0,t]} e^{-2\lambda s} (f'(s))^2 d[\phi]_s \]
([18], Chapter II, Theorem 29), and the boundedness of \( f' \) and (3) imply that \( Z \) is a square-integrable martingale (see also [18], Chapter II, Corollary 3 to Theorem 27). Then by the Optional Sampling Theorem for bounded stopping times we have \( E[Z_{T_n}] = 0 \). The boundedness of \( f' \) implies that the stochastic integral with respect to the Brownian motion also has zero expected value. On the other hand, \( U_{T(n)} \) has zero expectation because \( (S_n, n \geq 1) \) is a localizing sequence for the local martingale \( U \). Therefore we can write the identity
\[
\begin{align*}
  f(x) &= E_x \left[ e^{-\lambda T_n} f(X_{T(n)}) \right] - E_x \left[ \int_{(0,T_n]} e^{-\lambda s} \{ Af(X_s) - \lambda f(X_s) \} ds \right] \\
  &\quad - E_x \left[ \sum_{s \leq T_n} e^{-\lambda s} \{ f(X_s) - f(X_s + \Delta Y_s) \} \right].
\end{align*}
\]
Using conditions (i), (ii) we arrive at
\[
\begin{align*}
  f(x) &\leq E_x \left[ e^{-\lambda T_n} f(X_{T(n)}) \right] + E_x \left[ \int_0^{T_n} e^{-\lambda s} \phi(X_s) ds + \sum_{\tau_j \leq T_n} e^{-\lambda \tau_j} g(\xi_j) \right].
\end{align*}
\]
There is a minor problem here since (i) implies the inequality \( Af(x) - \lambda f(x) + \phi(x) \geq 0 \) only for \( x \in \mathbb{R} \setminus S \). However, if \( \sigma \neq 0 \) then (30) still holds by Lemma Appendix A.3 in Appendix A. If \( \sigma = 0 \) then the function \( x \mapsto Af(x) - \lambda f(x) + \phi(x) \) is right-continuous. Indeed, \( x \mapsto Af(x) \) is continuous because \( \sigma = 0 \) removes the discontinuity of the function \( x \mapsto Af(x) \) caused by the discontinuity of the second derivative of \( f \), and \( \phi \) is assumed to be right-continuous. Let \( \tilde{x} \in S \) arbitrary. Since \( S \) is finite, there exists a sequence \( (x_n) \subset \mathbb{R} \) such that \( x_n \downarrow \tilde{x} \) as \( n \to \infty \), and \( x_n \notin S \) for every \( n \). Then \( Af(x_n) - \lambda f(x_n) + \phi(x_n) \geq 0 \) and the right-continuity implies that the same inequality holds for \( x_n \) replaced by \( \tilde{x} \), which implies that inequality (i) holds for every \( x \in \mathbb{R} \). Next we take limits of both sides in (30) as \( n \to \infty \). Since \( f \) is continuously differentiable and has bounded derivative on \( \mathbb{R} \), it satisfies a linear growth condition \( f(x) \leq K_1 |x| + K_2 \) for some positive constants \( K_1, K_2 \), hence by (5)
\[
0 \leq e^{-\lambda t} f(X_t) \leq e^{-\lambda t} (K_1 |X_t| + K_2) \leq K_3.
\]
The first term on the right-hand side of (30) converges to \( E_x \left[ e^{-\lambda t} f(X_t) \right] \) by (31) and by the Bounded Convergence Theorem. For the second term on the right-hand side we apply the Monotone Convergence Theorem, and conclude that
\[
\begin{align*}
f(x) &\leq E_x \left[ e^{-\lambda T_n} f(X_{T(n)}) \right] + E_x \left[ \int_0^t e^{-\lambda s} \phi(X_s) ds + \sum_{\tau_j \leq t} e^{-\lambda \tau_j} g(\xi_j) \right].
\end{align*}
\]
Next we take the \( \lim \inf \) in (32) as \( t \to \infty \). By (10) the \( \lim \inf \) of the first term on the right-hand side of (32) is zero. We again apply the Monotone Convergence Theorem for the second term, and conclude that
\[
\begin{align*}
f(x) &\leq E_x \left[ \int_0^\infty e^{-\lambda s} \phi(X_s) ds + \sum_{j=1}^\infty e^{-\lambda \tau_j} g(\xi_j) \right],
\end{align*}
\]
which completes the proof.

**Proof of Theorem 3.2:** By Lemma 3.3 the conditions of Theorem 3.2 are stronger than those of Theorem 3.1, hence all statements in the proof of Theorem 3.1 are valid. In particular, (29) is still true. From condition (a), Lemma Appendix A.3, (19), and (20) follows that we have equalities in (30), (32), and (33).

4. Analysis of the Optimal Control \((T^*, \Xi^*)\)

We now set out to determine the transient and steady-state behavior of the controlled cash on hand process \(X_t\) under the optimal control \((T^*, \Xi^*)\) assuming that \(Y_1\) is a spectrally positive Lévy process. In other words,

\[
\int_{\mathbb{R}^-} \nu(dy) = 0
\]

and \(Y_t\) is not a subordinator. The case of a spectrally negative Lévy process may be treated similarly. Our main result will be to determine the Laplace transform (with respect to time) of the transition probabilities of \(X_t\) and also to determine the limiting distribution of \(X_t\).

We begin by determining the Laplace transform of the transition probabilities of \(X_t\). Let \(A \in \mathcal{B}(\mathbb{R})\) be a Borel set of \(\mathbb{R}\) and let \(e_q\) be an exponential random variable with rate \(q\) independent of \(X\). Now consider \(P_x[X_t \in A]\) for \(x \in (a, b)\). It then follows conditioning on the value of \(e_q\) relative to the stopping time \(\tau^*_1\) that

\[
P_x[X_t \in A] = E_x[1\{X_{\tau^*_1} \in A\}1\{e_q < \tau^*_1\}] + E_x[1\{X_{\tau^*_1} \in A\}1\{e_q \geq \tau^*_1\}].
\]

(34)

However, since \(X_t = Y_t\) for \(0 \leq t < \tau^*_1\), it follows by the memoryless property of the exponential distribution and the strong Markov property that \(E_x[1\{X_{\tau^*_1} \in A\}1\{e_q \geq \tau^*_1\}]\) is equal to

\[
E_x[1\{X_{\tau^*_1} \in A\}]P_x[\{e_q \geq \tau^*_1\} \cap \{Y(\tau^*_1) \leq a\}] + E_x[1\{X_{\tau^*_1} \in A\}]P_x[\{e_q \geq \tau^*_1\} \cap \{Y(\tau^*_1) \geq b\}]
\]

Substituting the above into (34), we have

\[
E_x[1\{X_{\tau^*_1} \in A\}] = E_x[1\{X_{\tau^*_1} \in A\}1\{e_q < \tau^*_1\}] + E_x[1\{X_{\tau^*_1} \in A\}]P_x[\{e_q \geq \tau^*_1\} \cap \{Y(\tau^*_1) \leq a\}] + E_x[1\{X_{\tau^*_1} \in A\}]P_x[\{e_q \geq \tau^*_1\} \cap \{Y(\tau^*_1) \geq b\}].
\]

(35)

It therefore remains to determine expressions for the three quantities on the righthand side of (35). We proceed term by term.

First note that \(\tau^*_1\) is equal to the first time the Lévy process \(Y_t\) exits the open interval \((a, b)\). We then have that in general one may write

\[
E_x[1\{X_{\tau^*_1} \in A\}1\{e_q < \tau^*_1\}] = E_x[1\{X_{\tau^*_1} \in A\}1\{e_q < \tau^*_1\}]
\]

(36)

\[
= q\int_0^{\infty} e^{-qt}E_x[1\{Y_t \in A\}1\{\tau^*_1 > t\}]dt
\]

\[
= q\int_A \int_0^{\infty} e^{-qt}P_x[\{Y(t) \in dy, \tau^*_1 > t\}]dt
\]

\[
= q\int_A U^{(q)}(x, dy),
\]

where \(U^{(q)}\) is the \(q\)-potential measure of \(Y_t\) (see (B.1)). This then provides an expression for the first term on the righthand side of (35). Also note, by Theorem 8.7 of [15] (quoted in this paper under Theorem Appendix B.2), if \(Y_t\) is spectrally positive then its \(q\)-potential measure \(U^{(q)}(x, dy)\) has a density \(u^{(q)}(x, y)\) given by

\[
u^{(q)}(x, y) = W^{(q)}(b - x)W^{(q)}(y - a) - W^{(q)}(y - x),
\]

(37)
whenever $s$ is large enough so that $\psi(-s) > q$ and $\psi(s) = \log E_0[e^{sY}]$ is the Laplace exponent of $Y_t$. Note that $\psi(s) < \infty$ for all $s \leq 0$ by the spectral positivity of $Y$. The function $W^{(0)}$ is called the scale function because it plays similar role to the scale function of a diffusion process; see Theorem Appendix B.1 (iii) in Appendix B.

Next note that setting $x = \alpha$ in (35), we obtain
\begin{equation}
E_\alpha[1]\{X_{e_\alpha} \in A\}[1 - P_\alpha[\{e_\alpha \geq \tau_1^x\} \cap \{Y(\tau_1^x) \leq a\}) - E_\beta[1]\{X_{e_\alpha} \in A\}]P_\alpha[\{e_\alpha \geq \tau_1^x\} \cap \{Y(\tau_1^x) \geq b\}]
\end{equation}
\begin{equation}
= E_\alpha[1]\{X_{e_\alpha} \in A\}1\{e_\alpha < \tau_1^x\},
\end{equation}
and similarly, setting $x = \beta$, we have
\begin{equation}
E_\alpha[1]\{X_{e_\alpha} \in A\}[1 - P_\beta[\{e_\beta \geq \tau_1^x\} \cap \{Y(\tau_1^x) \geq b\}]) - E_\alpha[1]\{X_{e_\alpha} \in A\}]P_\beta[\{e_\beta \geq \tau_1^x\} \cap \{Y(\tau_1^x) \geq a\}]
\end{equation}
\begin{equation}
= E_\alpha[1]\{X_{e_\alpha} \in A\}1\{e_\alpha < \tau_1^x\},
\end{equation}

(39) and (40) constitute a set of linear equations for $E_\alpha[1\{X_{e_\alpha} \in A\}]$ and $E_\beta[1\{X_{e_\alpha} \in A\}]$. Moreover, so long as $q > 0$, we have that
\begin{equation}
(1 - P_\beta[\{e_\beta \geq \tau_1^x\} \cap \{Y(\tau_1^x) \geq b\}]) - P_\alpha[\{e_\alpha \geq \tau_1^x\} \cap \{Y(\tau_1^x) \geq b\}])
\end{equation}

and so the determinant associated with (39) and (40) is non-zero and hence a solution exists. Solving for $E_\alpha[1\{X_{e_\alpha} \in A\}]$ and $E_\beta[1\{X_{e_\alpha} \in A\}]$ then yields that $E_\alpha[1\{X_{e_\alpha} \in A\}]$ is given by
\begin{equation}
\frac{1}{C_{\alpha,\beta,b}} \times \left( E_\alpha[1\{X_{e_\alpha} \in A\}1\{e_\alpha < \tau_1^x\}] - P_\beta[\{e_\beta \geq \tau_1^x\} \cap \{Y(\tau_1^x) \geq b\}])
\end{equation}
and $E_\beta[1\{X_{e_\alpha} \in A\}]$ is given by
\begin{equation}
\frac{1}{C_{\alpha,\beta,b}} \times \left( E_\beta[1\{X_{e_\alpha} \in A\}1\{e_\beta < \tau_1^x\}] - P_\alpha[\{e_\alpha \geq \tau_1^x\} \cap \{Y(\tau_1^x) \leq a\}]
\end{equation}
where
\begin{equation}
C_{\alpha,\beta,b} = (1 - P_\beta[\{e_\beta \geq \tau_1^x\} \cap \{Y(\tau_1^x) \geq b\}]) - P_\alpha[\{e_\alpha \geq \tau_1^x\} \cap \{Y(\tau_1^x) \leq a\}]
\end{equation}

We now proceed to compute the terms appearing on the right-hand sides of (41) and (42).

First note that the expressions of the form $E_x[1\{X_{e_\alpha} \in A\}1\{e_\alpha < \tau_1^x\}]$ appearing in (41) and (42) have already been determined by (36). It remains to determine expressions for the probabilities appearing in (41) and (42). However, note that
\begin{equation}
P_x[\{e_\alpha \geq \tau_1^x\} \cap \{Y(\tau_1^x) \leq a\}] = \int_0^\infty q e^{-sY}P_x[\{t \geq \tau_1^x\} \cap \{Y(\tau_1^x) \leq a\}]dt = E_x\left[\int_{\tau_1^x}^\infty q e^{-sY}dt\right]1\{Y(\tau_1^x) \leq a\]
where the final equality follows from Theorem 8.1 of [15] (quoted in this paper under Theorem Appendix B.1 in Appendix B) and we have the relationship

\[ Z^{(q)}(x) = 1 + q \int_0^x W^{(q)}(y) dy. \]

In a similar fashion, using Theorem 8.1 of [15] (quoted in this paper under Theorem Appendix B.1 in Appendix B), one may compute

\[ P_x[\{ e_q \geq \tau_1 \} \cap \{ Y(\tau_1) \geq b \}] = \frac{W^{(q)}(b-x)}{W^{(q)}(b-a)}. \]

Substituting the above results into (41) and (42), and subsequently into (35), one obtains an expression for the limiting distribution of \( X_t \) as \( t \to \infty \). Our main result in this regard will be the following. Recall the definition of \( U^{(q)} \) as the \( q \)-potential measure associated with \( Y \).

**Proposition 4.1.** If \( Y_t \) is spectrally positive, then under a double bandwidth control policy \((a, \alpha, \beta, b)\), the limiting distribution of \( X_t \) is given by

\[ \pi(A) = \frac{1}{K_{a,\alpha,\beta,b}} \left( \left( 1 - \frac{U^{(0)}(b-\beta)}{U^{(0)}(b-a)} \right) \int_A U^{(0)}(\alpha, dy) + \left( \frac{W^{(0)}(b-a)}{W^{(0)}(b-a)} \right) \int_A U^{(0)}(\beta, dy) \right), \]

for each \( A \in \mathcal{B} \), where \( K_{a,\alpha,\beta,b} \) is an appropriate normalizing constant, as given by (45).

We prove Proposition 4.1 in a series of lemmas. First note that by the strong Markov property, \( X_t \) is a regenerative process with possible regeneration points either \( \alpha \) or \( \beta \). Let us consider the point \( \alpha \) and define \( n^*_\alpha = \inf\{ n \geq 1 : X(\tau^*_n) = \alpha \} \). By the standard theory of regenerative processes, see for instance Theorem 1.2 of Chapter VI of [1], if we may show that \( E_\alpha[\tau^*_n] \) is nonlattice\(^1\), then \( \lim_{t \to \infty} P_x(X_t \in A) = \pi(A) \) exists for all \( A \in \mathcal{B}(\mathbb{R}) \) and \( x \in \mathbb{R} \) and is given by

\[ \pi(A) = \frac{E_\alpha[\int_0^{\tau^*_n} 1\{ X_s \in A \} ds]}{E_\alpha[\tau^*_n]} . \]

The following lemma now shows that \( E_\alpha[\tau^*_n] \) is nonlattice.

**Lemma 4.2.** If the Lévy process \( Y_t \) is spectrally positive, then \( E_\alpha[\tau^*_n] < \infty \).

**Proof:** Note first that

\[ \tau^*_n = \tau^*_1 1\{ Y(\tau^*_1) \leq a \} + \tau^*_n 1\{ Y(\tau^*_n) \geq b \} = \tau^*_1 + (\tau^*_n - \tau^*_1) 1\{ Y(\tau^*_1) \geq b \} . \]

Hence, by the strong Markov property,

\[ E_\alpha[\tau^*_n] = E_\alpha[\tau^*_1] + E_\alpha[(\tau^*_n - \tau^*_1) 1\{ Y(\tau^*_1) \geq b \}] \]

---

\(^1\)A random variable is called **lattice** if its distribution is concentrated on a set of the form \( \{ \delta, 2\delta, \ldots \} \). It is called nonlattice if it is not.
\[ E_\alpha[\tau_1^*] + E_\beta[\tau_n^*]P_\alpha[Y(\tau_1^*) \geq b]. \]

Similarly, we may show

\[ E_\beta[\tau_n^*] = E_\beta[\tau_1^*] + E_\beta[\tau_n^*]P_\beta[Y(\tau_1^*) \geq b], \]

from which we obtain

\[ E_\beta[\tau_n^*] = \frac{E_\beta[\tau_1^*]}{1 - P_\beta[Y(\tau_1^*) \geq b]}. \]

Now note that since \( Y_t \) is spectrally positive, we have by Theorem 8.1 of [15] (quoted in this paper under Theorem Appendix B.1 in Appendix B) that

\[ P_\beta[Y(\tau_1^*) \geq b] = \frac{W(0)(b - \beta)}{W(0)(b - a)} < 1 \]

and so it suffices from the above to show \( E_\alpha[\tau_1^*], E_\beta[\tau_1^*] < \infty \).

We now show that in general for \( x \in (a, b) \),

\[ E_x[\tau_*] < \infty. \]

Recall by [15], the potential measure of \( Y_t \) upon exiting \([a, b]\) is given by

\[ U(x, dy) = \int_0^\infty P_x[Y_t \in dy, \tau_* > t]dt. \]

Integrating over \([a, b]\), we obtain that

\[ \int_{[a,b]} U(x, dy) = \int_{[a,b]} \int_0^\infty P_x[Y_t \in dy, \tau_* > t]dt \]

\[ = \int_0^\infty \int_{[a,b]} P_x[Y_t \in dy, \tau_* > t]dt \]

\[ = \int_0^\infty P_x[\tau_* > t]dt \]

\[ = E_x[\tau_*]. \]

However, by Theorem 8.7 of [15] (quoted in this paper in Appendix B, Theorem Appendix B.2), since \( Y_t \) is spectrally positive, \( U(x, dy) \) has a density given by

\[ u(x, y) = \frac{W(0)(b - x)}{W(0)(b - a)} - W(0)(y - x). \]

Integrating over \([a, b]\), we therefore find that

\[ \int_{[a,b]} U(x, dy) = \int_{[a,b]} \left( \frac{W(0)(b - x)}{W(0)(b - a)} - W(0)(y - x) \right) dy \]

\[ < \infty, \]

where the inequality follows since \( W(0) \), due to its continuity on \([0, \infty)\) (see [15] Theorem 8.1, quoted in this paper in Appendix B, Theorem Appendix B.1), is bounded on compact sets. By the above, this completes the proof. \( \bullet \)

The following lemma allows us to take the limit as \( q \to 0 \) in (41) and (42) in order to obtain the limiting distribution of \( X_t \).

**Lemma 4.3.** For each \( x \in \mathbb{R} \),

\[ \lim_{q \to 0} P_x[X(e_q) \in A] = \pi(A) \]
Thus, noting that we see that both the numerator and denominator in (41) are differentiable.

\[
|P_x[X_\tau \in A] - \pi(A)| = \left| \int_0^\infty \{P_x[X_\tau \in A] - \pi(A)\} \, dq \right|
\]

\[
\leq \left| \int_0^T \{P_x[X_\tau \in A] - \pi(A)\} \, dq \right| + \left| \int_T^\infty \{P_x[X_\tau \in A] - \pi(A)\} \, dq \right|.
\]

The second term in the above expression is bounded by \(\epsilon\) and the first term converges to zero as \(q \to 0\) by the Dominated Convergence Theorem, which completes the proof.

**Proof of Proposition 4.1:** Using Lemma 4.3, we now wish to take limits \(q \to 0\) in (41) in order to determine the limiting distribution \(\delta\). However, both the numerator and denominator in (41) converge to 0 as \(q \to 0\) and so we must apply L'Hopital's rule. Before doing so, however, we first must verify that both the numerator and denominator in (41) are differentiable.

By Lemma 8.3 and Corollary 8.5 in [15] we have that for each \(x > 0\), both \(W^q(x)\) and \(Z^q(x)\) are differentiable in \(q\). Moreover, since

\[
\frac{W^q(b-x)}{W^q(b-a)} = E_x[e^{-q\tau_1} 1\{Y_{\tau_1} \geq b\}],
\]

for \(a \leq x < b\), it follows that

\[
\frac{d}{dq} \frac{W^q(b-x)}{W^q(b-a)} = -E_x[qe^{-q\tau_1} 1\{Y_{\tau_1} \geq b\}] < \infty,
\]

where the inequality follows as in the proof of Lemma 4.2. Finally, since for each \(a \leq x \leq b\), \(U^q(x,dy)\) has a density \(u^q(x,y)\) given by (37) it follows that for each \(A \in B(\mathbb{R})\),

\[
\frac{d}{dq} \int_A U^q(x,dy) = \int_A \frac{d}{dq} u^q(x,y)dy.
\]

Thus, noting that

\[
(E_x[1\{X_{e_q} \in A\}1\{e_q < \tau_1\}](1 - P_\delta[\{e_q \geq \tau_1\} \cap \{Y(\tau_1) \geq b\}])
+ E_x[1\{X_{e_q} \in A\}1\{e_q < \tau_1\}]P_\delta[\{e_q \geq \tau_1\} \cap \{Y(\tau_1) \geq b\}])
= q \int_A U^q(\alpha,dy) \left(1 - \frac{W^q(b-\beta)}{W^q(b-a)}\right) + q \int_A U^q(\beta,dy) \left(\frac{W^q(b-\alpha)}{W^q(b-a)}\right)
\]

and

\[
((1 - P_\delta[\{e_q \geq \tau_1\} \cap \{Y(\tau_1) \geq b\}])P_\delta[\{e_q \geq \tau_1\} \cap \{Y(\tau_1) \leq a\}]
- P_\delta[\{e_q \geq \tau_1\} \cap \{Y(\tau_1) \geq b\}]P_\delta[\{e_q \geq \tau_1\} \cap \{Y(\tau_1) \leq a\}])
= \left(1 - \frac{W^q(b-\beta)}{W^q(b-a)}\right) \left(1 - \frac{W^q(b-\alpha)}{W^q(b-a)}\right)
- \left(\frac{Z^q(b-\beta)}{W^q(b-a)}\right) \left(\frac{W^q(b-\alpha)}{W^q(b-a)}\right),
\]

we see that both the numerator and denominator in (41) are differentiable.

Let us now take derivatives on the righthand sides of (43) and (44).
Taking the derivative of the right hand side of (43) and evaluating at \( q = 0 \) we obtain
\[
\int_{\mathcal{A}} U^{(\nu)}(\alpha, dy) \left( 1 - \frac{W^{(\nu)}(b - \beta)}{W^{(\nu)}(b - a)} \right) + \int_{\mathcal{A}} U^{(\nu)}(\beta, dy) \left( \frac{W^{(\nu)}(b - \alpha)}{W^{(\nu)}(b - a)} \right).
\]

Next, recalling that
\[
Z^q(x) = 1 + q \int_0^x W^q(y)dy,
\]
it follows upon taking the derivative of the righthand side of (44) and evaluating at \( q = 0 \) that we obtain
\[
\frac{d}{dq} \left( \left( 1 - \frac{W^{(\nu)}(b - \beta)}{W^{(\nu)}(b - a)} \right) \left( 1 - \frac{Z^{(\nu)}(b - \alpha) - Z^{(\nu)}(b - a)}{W^{(\nu)}(b - a)} \right) \right)
\]
\[
= W^{(\nu)}(b - \beta) \int_0^{b - \alpha} W^q(x)dx + \frac{W^{(\nu)}(b - \alpha)}{W^{(\nu)}(b - a)} \int_{b - \beta}^{b - \alpha} W^q(x)dx - \int_0^{b - \alpha} W^q(x)dx
\]
\[
= K_{a,\alpha,\beta,b}.
\]

Thus, by (41) and Lemma 4.2 we obtain the desired result. 

5. An Example

We suppose in this section that
\[
Y_t = x + \sigma w_t + N_t,
\]
where \( N \) is a compound Poisson process independent of \( w \) such that the rate of jump arrivals is equal to 1 and the Lévy measure \( \nu \) of \( N \) is
\[
\nu(dy) = \theta e^{-\theta y}dy, \quad y \geq 0
\]
for some \( \theta > 0 \), and \( \nu((\infty, 0]) = 0 \). In this case, we can write the linear operator \( \mathcal{A} \) in the form
\[
\mathcal{A}f(x) = \frac{\sigma^2}{2} f''(x) + \int_0^\infty [f(x + y) - f(x)] \nu(dy).
\]

We also specify the opportunity cost function as
\[
\phi(x) = (x - \rho)^2
\]
where \( \rho \) is a fixed target value. Suppose now that \( x \in (a, b) \). Using (46), the equation in (i) in Theorem 3.1 may be written as
\[
\frac{1}{2} \sigma^2 f''(x) + (x - \rho)^2 - \lambda f(x) + \int_0^\infty [f(x + y) - f(x)] \theta e^{-\theta y}dy = 0.
\]

This becomes
\[
\frac{1}{2} \sigma^2 f''(x) + (x - \rho)^2 - (1 + \lambda) f(x) + \theta e^{\theta x} \int_x^{b} f(z)e^{-\theta z}dz + e^{\theta x} \int_0^\infty [f(b) + d(z - b)] \theta e^{-\theta z}dz = 0,
\]
and also
\[
\frac{1}{2} \sigma^2 e^{-\theta x} f''(x) + e^{-\theta z}(x - \rho)^2 - (1 + \lambda)e^{-\theta x} f(x) + \theta \int_x^{b} f(z)e^{-\theta z}dz + \zeta = 0,
\]

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Let us now introduce $e^{-\theta x} f(x) = g(x)$. We then obtain the following equation from (48):

$$
\left( \frac{1}{2} \sigma^2 \theta^2 - \lambda - 1 \right) g(x) + \sigma^2 \theta g'(x) + \frac{1}{2} \sigma^2 \theta^2 (x - \rho)^2 + \theta \int_x^b g(z) dz + \zeta = 0.
$$

(49)

Differentiating the above with respect to $x$ we get the following inhomogeneous linear ordinary differential equation of the third order:

$$
\frac{1}{2} \sigma^2 g'''(x) + \sigma^2 \theta g''(x) + \left( \frac{1}{2} \sigma^2 \theta^2 - \lambda - 1 \right) g'(x) - \theta g(x) + 2e^{-\theta x}(x - \rho) - \theta e^{-\theta x}(x - \rho)^2 = 0.
$$

(50)

A particular solution for the inhomogeneous equation, denoted by $g_p$, is given by

$$
g_p(x) = e^{-\theta x} \left[ K_1 (x - \rho)^2 + K_2 (x - \rho) + K_3 \right],
$$

where

$$
K_1 = \frac{1}{\lambda}, \quad K_2 = \frac{2}{\theta \lambda^2}, \quad K_3 = \frac{1}{\lambda^3 \theta^2} [2\lambda + 2 + \theta^2 \lambda \sigma^2].
$$

The general solution of the homogeneous equation is given by $g_h$, that is,

$$
g_h(x) = L_1 e^{c_1 x} + L_2 e^{c_2 x} + L_3 e^{c_3 x}
$$

where $c_1, c_2, c_3$ are the roots of the equation

$$
P(x) = \frac{1}{2} \sigma^2 x^3 + \sigma^2 \theta x^2 + \left( \frac{1}{2} \sigma^2 \theta^2 - \lambda - 1 \right) x - \theta = 0
$$

and $L_1, L_2, L_3$ are “free parameters”. Notice that $P(0) = -\theta < 0$, $P(-\theta) = \theta \lambda > 0$, and $\lim_{x \to -\infty} P(x) = \infty$, thus $P(x)$ has three roots, say $c_1 < -\theta$, $-\theta < c_2 < 0$ and $c_3 > 0$.

We have now arrived at the following family of candidate solutions:

$$
g(x; L_1, L_2, L_3) = e^{-\theta x} \left[ K_1 (x - \rho)^2 + K_2 (x - \rho) + K_3 \right] + L_1 e^{c_1 x} + L_2 e^{c_2 x} + L_3 e^{c_3 x}.
$$

This gives

$$
f(x; L_1, L_2, L_3) = K_1 (x - \rho)^2 + K_2 (x - \rho) + K_3 + L_1 e^{(\theta + c_1) x} + L_2 e^{(\theta + c_2) x} + L_3 e^{(\theta + c_3) x}.
$$

For simplicity we shall use the notation $f(x; L_1, L_2, L_3, b) = f(x)$.

We now have 7 unknown parameters $a, \alpha, \beta, b, L_1, L_2, L_3$. From the conditions of Theorem 3.1, we may derive the following 6 equations for these constants:

$$
f'(a) = -c
$$

(51)

$$
f'(\alpha) = -c
$$

(52)

$$
f'(b) = d
$$

(53)

$$
f'(\beta) = d
$$

(54)
In addition, if we trace back our derivation in the above, then we see that we must have (47) hold for at least for one particular \( x \) since in going from (49) to (50) we took a derivative. Select \( x = b \). This then gives us our 7th equation

\[
\frac{1}{2} \sigma^2 f''(b) + (b - \rho)^2 - (1 + \lambda)f(b) + \epsilon \theta \zeta = 0.
\]

(56)

We now have the following.

**Theorem 5.1.** Suppose that there exist seven constants \( L_1 \leq 0, L_2 \leq 0, L_3 \leq 0, a < \alpha \leq \beta < b \) satisfying the seven equations (51)-(55) and (56). We define \( h \) by

\[
h(x) = \begin{cases} 
  f(a) - c(x - a), & \text{if } x \leq a, \\
  f(x), & \text{if } a \leq x \leq b \\
  f(b) + d(x - b), & \text{if } x \geq b.
\end{cases}
\]

Then \( h(x) = V(x) \), i.e., \( h(x) \) is the value function of the optimization problem. Furthermore, the policy \((T^*, \Xi^*)\) described in (14) and (16) with this choice of \( a, \alpha, \beta, b \) is optimal.

**Example.** We now provide an example where the 7 equations above may be numerically solved for. Note that this implies that the equations are not vacuous. Suppose that \( \rho = 0, \theta = 1, \sigma = 1, \lambda = 1/64, C = 1/25, c = 0, D = 1/50 \) and \( d = 7/10 \). Then, using Mathematica the 7 constants of Theorem 5.1 may found up to a high level of precision. In particular, up to 2 decimals they are given by \( a = -1.13, \alpha = -0.56, \beta = 0.37, b = 0.84, L_1 = -1.77, L_2 = -536504.65, L_3 = -0.05 \).

In order to prove Theorem 5.1, we need the following lemma.

**Lemma 5.1.** Assume the conditions of Theorem 5.1. Then there exists a constant \( \xi \in (\alpha, \beta) \) such that \( h' \) is convex on \([a, \xi]\), concave on \([\xi, b]\). Furthermore \( h'(x) \leq -c \) if \( x \in [a, \alpha] \), \( h'(x) \geq d \) if \( x \in [\beta, b] \), and \(-c \leq h'(x) \leq d \) if \( x \in [\alpha, \beta] \).

**Proof:** From the condition that \( L_1, L_2, L_3 \leq 0 \) it follows that \( h'''(x) \) is decreasing on \((a, b)\). Therefore \( h''(x) \) has at most two zero points, which implies that \( h'(x) \) has at most two local extreme values in \((a, b)\). By (51)-(53) the derivative function \( h' \) must be first decreasing then increasing then again decreasing on \([a, b]\). Since \( h'' \) is concave, it must be either increasing, or decreasing, or first increasing then decreasing on \([a, b]\). But the first two possibilities are not possible, since \( h' \) can not be neither convex nor concave on \([a, b]\). Therefore it must be first convex then concave.

**Proof of Theorem 5.1:** We need to prove that the conditions of Theorem 3.1 are satisfied. Condition (i) and the required smoothness of \( h \) follows from our construction. Next we prove (ii) and (iv). From conditions (52), (53) and Lemma 5.1 it follows that

\[
Mh(x) = \begin{cases} 
  h(\alpha) + C + c(\alpha - x), & \text{if } a \leq x \leq \alpha, \\
  h(x) + \min\{C, D\}, & \text{if } \alpha < x < \beta, \\
  h(\beta) + D + d(x - \beta), & \text{if } \beta \leq x \leq b.
\end{cases}
\]

Condition (ii) follows from (54) and (55). We show condition (iv) for the case of \( x \in [a, \alpha] \), the case of \( x \in [\beta, b] \) is similar and the case of \( x \in (\alpha, \beta) \) is obvious. We need to show that \( 0 \leq h(\alpha) - h(x) + C + c(\alpha - x) \). For \( x = a \) we have equality by (54), and the derivative of the right-hand side of the inequality with respect to \( x \) is \(-h'(x) - c\), which is non-negative on \([a, \alpha]\) by Lemma 5.1. Next we show condition (iii). First we look at the case of \( x > b \). Let

\[
K(x) = Ah(x) - \lambda h(x) + (x - \rho)^2;
\]

\( x \in \mathbb{R} \setminus \{a, b\} \).
\[
K(x) = \frac{\sigma^2}{2}h''(x) + \int_0^\infty [h(x + y) - h(x)] \nu(dy) - \lambda h(x) + (x - \rho)^2 = 0, \quad x \in (a, b)
\]  
(57)

and
\[
K(x) = \int_0^\infty [h(x + y) - h(x)] \nu(dy) - \lambda h(x) + (x - \rho)^2, \quad x \in (-\infty, a) \cup (b, \infty).
\]  
(58)

Therefore \(0 = K(b-) = \frac{\sigma^2}{2}h''(b-) + K(b+),\) and \(\frac{\sigma^2}{2}h''(b-) \leq 0\) implies \(K(b+) \geq 0.\) On the other hand for \(x > b\) we have \(K'(x) = -\lambda d + 2(x - \rho)\) and \(K'(b+) = -\lambda d + 2(b - \rho).\) We also have by (57)
\[
0 = K'(x) = \frac{\sigma^2}{2}h'''(x) + \int_0^\infty [h'(x + y) - h'(x)] \nu(dy) - \lambda h'(x) + 2(x - \rho), \quad x \in (a, b)
\]  
(59)

hence
\[
0 = K'(b-) = \frac{\sigma^2}{2}h'''(b-) - \lambda d + 2(b - \rho) = \frac{\sigma^2}{2}h'''(b-) + K'(b+).
\]

Since \(h'''(b-) \leq 0\) so we must have \(K'(b+) \geq 0.\) But \(K'\) is increasing on \((b, \infty),\) thus \(K'(x) \geq 0\) whenever \(x \in (b, \infty).\) This in turn implies that \(K\) is increasing on \((b, \infty),\) thus \(K(x) \geq 0\) for \(x \in (b, \infty).\)

Next we show that \(K(x) \geq 0\) for \(x < a.\) By (57) and (58) \(0 = K(a+) = K(a-)+\frac{\sigma^2}{2}h'''(a+)\) which implies that \(K(a-) \geq 0.\) Hence all we need to show is that \(K'(x) \leq 0\) for \(x \leq a.\) Differentiating (58) we get
\[
K'(x) = \int_0^\infty [h'(x + y) - h'(x)] \nu(dy) - \lambda h'(x) + 2(x - \rho), \quad x < a,
\]

and with a change of variable in the integral one can see that
\[
K'(x) = e^{\theta(x-a)} \int_0^\infty [h'(a + z) + c] \nu(dz) + \lambda c + 2(x - \rho), \quad x < a
\]
(60)

and
\[
K''(x) = \theta e^{\theta(x-a)} \int_0^\infty [h'(a + z) + c] \nu(dz) + 2, \quad x < a.
\]
(61)

Thus \(K''\) is either increasing or decreasing on \((-\infty, a)\) depending on the sign of \(\int_0^\infty [h'(a + z) + c] \nu(dz)\) which makes \(K'\) either convex or concave on \((-\infty, a)\). However, the fact that \(\lim_{x \to -\infty} K'(x) = -\infty\) implies that \(K'\) must be concave and \(K''\) decreasing on \((-\infty, a)\). In addition, the integral in (61) and in (60) is negative. By (59) and (60)
\[
0 = K'(a+) = \frac{\sigma^2}{2}h'''(a+) + K'(a-)
\]

thus \(K'(a-) \leq 0\) follows from \(h'''(a+) \geq 0.\) Formula (60) imply that \(K'\) has at most one zero-point on \((-\infty, a)\). These facts about \(K'\) imply that indeed \(K'(x) \leq 0.\) •

We conclude this section by providing an example showing how the limiting distribution \(\pi\) of Proposition 4.1 may be calculated for an arbitrary double bandwidth control policy \((a, \alpha, \beta, b)\). We will assume that \(Y_t = \vartheta t + N_t,\) where \(\vartheta < 0\) and \(N_t\) is a compound Poisson process which has jumps at rate one and jump sizes which are exponentially distributed with rate \(\theta.\) Note that by (37) and Proposition 4.1, it suffices to determine the function \(W^{(0)} = \lim_{q \to 0} W^{(q)}\). By (38), the Laplace transform of \(W^{(0)}\) is given by \(1/\psi(-s)\) where \(\psi(s)\) is the Lévy exponent of \(Y_t.\) Moreover, by (8.1) in [15] it then follows that
\[
\psi(s) = \vartheta s - \int_0^\infty (1 - e^{sx}) \theta e^{-\theta x} dx,
\]
for \(s < \theta\), which reduces to \(\psi(s) = \vartheta s + s(\theta - s)^{-1}.\) One may now proceed to verify that
\[
\frac{1}{\psi(-s)} = \frac{-\theta}{s(\vartheta s + \vartheta \theta + 1)} - \frac{1}{\vartheta s + \vartheta \theta + 1}.
\]

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In the case in which $\theta \vartheta \neq -1$ inverting each of the terms in the above, one obtains that the function $W^{(0)}$ is given by

$$W^{(0)}(x) = \frac{-\theta}{\theta \vartheta + 1} - \left(\frac{1}{\vartheta} - \frac{\theta}{\theta \vartheta + 1}\right) \exp\left(-\frac{\theta \vartheta + 1}{\vartheta} x\right).$$

For the case in which $\theta \vartheta = 1$, one has that $W^{(0)}(x) = \vartheta^{-1}(\theta x + 1)$. Substituting into the formula of Proposition 4.1, one may now obtain the density of $\pi$.

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Appendix A. Appendix

In Appendix A, we provide a proof of the fact that Itô’s rule applies to test functions $f \in \mathcal{D}$. For $f \in \mathcal{D}$ the second derivative $f''(x)$ may not exist in points $S = \{x_1, \ldots, x_m\}$. We shall call $S$ the set of exceptional points. We extend $f''$ to the entire of $\mathbb{R}$ by assuming an arbitrary value for $f''(x_i)$. This convention will be used in the rest of this section. The following is then the main result.

**Proposition Appendix A.1.** If $f \in \mathcal{D}$ and $X$ is a controlled cash on hand process with an arbitrary impulse control $(T, \Xi) = (\tau_1, \tau_2, \ldots, \tau_n, \xi_1, \xi_2, \ldots, \xi_n, \ldots)$ then Itô’s rule holds in its usual form:

$$f(X_t) - f(X_0) = \int_{(0,t]} f'(X_{s-})dX_s + \frac{\sigma^2}{2} \int_{(0,t]} f''(X_s)ds + \sum_{0 < s \leq t} \{f(X_s) - f(X_{s-}) - f'(X_{s-})\Delta X_s\}. \quad (A.1)$$

In order to prove Proposition Appendix A.1, we need the following two lemmas.

**Lemma Appendix A.2.** Let $f \in \mathcal{D}$ with set of exceptional points $S$. Then there exists a sequence $(f_n)_{n \geq 1} \subset C^2(\mathbb{R})$ such that the following hold:

(i) $f_n(x) \to f(x)$ and $f_n'(x) \to f'(x)$ for every $x \in \mathbb{R}$ as $n \to \infty$;

(ii) $f_n''(x) \to f''(x)$ for every $x \in \mathbb{R} \setminus S$ as $n \to \infty$;

(iii) $f_n'$ and $f_n''$ are bounded uniformly in $n$, i.e., $|f_n'(x)| \leq C_1$ and $|f_n''(x)| \leq C_1$ for some constant $C_1$ and all $n$ and $x \in \mathbb{R}$.

The proof of this lemma can be based on the proof of a similar lemma in Øksendal [16], Appendix D with some obvious modifications.

**Lemma Appendix A.3.** If $\sigma \neq 0$ then for every $z \in \mathbb{R}$

$$\int_0^\infty 1_{\{z\}}(X_s)ds = 0, \quad P_x-\text{a.s.}$$

In other words, the Lebesgue measure of the time the controlled cash on hand process spends at level $z$ is zero.

**Proof:**

$$E_x \left[\int_0^\infty 1_{\{z\}}(X_s)ds\right]$$
We deal with these last two integrals separately.

\[ \int_0^\infty P_x[\tau_i = s] ds \leq \int_0^\infty P_x[\tau_i = s] ds + \int_0^\infty P_x[\tau_i = s, \tau_i \neq s] ds. \]

We then have

\[ M \]

and this last expression is zero because the cardinality of the set \{s \geq 0 : P_x[\tau_i = s] > 0\} is either countably infinite or finite. For the second integral we have

\[ \int_0^\infty P_x[\tau_i = s, \tau_i \neq s] ds = \sum_{i=1}^\infty \int_0^\infty P_x[\tau_i = s] ds = \sum_{i=1}^\infty \int_0^\infty P_x[\tau_i = s] ds \]

and this last expression is zero because the cardinality of the set \{s \geq 0 : P_x[\tau_i = s] > 0\} is either countably infinite or finite. For the second integral we have

\[ \int_0^\infty P_x[\tau_i = s, \tau_i \neq s] ds = \sum_{i=1}^\infty \int_0^\infty P_x[\tau_i = s] ds \]

and this last expression is zero because the cardinality of the set \{s \geq 0 : P_x[\tau_i = s] > 0\} is either countably infinite or finite. For the second integral we have

\[ \int_0^\infty P_x[\tau_i = s, \tau_i \neq s] ds = \sum_{i=1}^\infty \int_0^\infty P_x[\tau_i = s] ds + \int_0^\infty P_x[\tau_i = s, \tau_i \neq s] ds. \]

Now it suffices to show that the probabilities in integrands on the right-hand side are zero. Indeed,

\[ P_x[X_s = z, \tau_i < s < \tau_{i+1}] \]

\[ = \int_{[0,s] \times \mathbb{R}} P_x[X_s = z, s < \tau_{i+1} | \tau_i = u, X_u = y] P_x[\tau_i = du, X_u = dy] \]

\[ = \int_{[0,s] \times \mathbb{R}} P_x[X_s - X_u = z - y, s < \tau_{i+1} | \tau_i = u, X_u = y] P_x[\tau_i = du, X_u = dy] \]

\[ = \int_{[0,s] \times \mathbb{R}} P_x[Y_s - Y_u = z - y, s < \tau_{i+1} | \tau_i = u, X_u = y] P_x[\tau_i = du, X_u = dy] \]

\[ \leq \int_{[0,s] \times \mathbb{R}} P_x[Y_s - Y_u = z - y, \tau_i = u, X_u = y] P_x[\tau_i = du, X_u = dy] \]

\[ = \int_{[0,s] \times \mathbb{R}} P_x[Y_s - Y_u = z - y] P_x[\tau_i = du, X_u = dy] \]

\[ = \int_{[0,s] \times \mathbb{R}} P_x[Y_{s-u} = z - y] P_x[\tau_i = du, X_u = dy] \]

and \( P_x[Y_{s-u} = z] = 0 \) follows from our assumption \( \sigma \neq 0 \) and Sato [20], Theorem 27.4. For the second probability on the right-hand side of (A.2) we have that \( P_x[X_s = z, s < \tau_i] \leq P_x[Y_s = z] = 0 \), by the same reference.

We now provide the proof of Proposition Appendix A.1.

**Proof of Proposition Appendix A.1:** Let \((f_n)_{n \geq 1}\) be the sequence approximating \(f\) in the sense of Lemma Appendix A.2. Itô's rule holds for each \(f_n\), i.e.,

\[ f_n(X_t) - f_n(X_0) = \int_{(0,t]} f''_n(X_s\cdot) dX_s + \frac{\sigma^2}{2} \int_{(0,t]} f''_n(X_s) ds \]

\[ + \sum_{0 \leq s \leq t} \left( f_n(X_s) - f_n(X_{s-}) - f''_n(X_{s-}) \Delta X_s \right). \]

All we need to show is that all three terms in the right-hand side of the above identity converge to the corresponding terms in the right-hand side of (A.1) as \( n \to \infty \). We can write \( X = X_0 + M_1(t) + A_1(t) \) where \( M_1 \) is a local martingale with bounded jumps (thus also locally square-integrable) and \( A_1 \) is a finite variation process (Jacod & Shiryaev [13], Proposition I.4.17). We then have

\[ \int_{(0,t]} f''_n(X_{s-}) dM_1(s) \to \int_{(0,t]} f''(X_{s-}) dM_1(s) \text{ in probability as } n \to \infty \]
by Theorem I.4.40 iii' in Jacod & Shiryaev [13]. Also
\[ \int_{[0, t]} f_n'(X_{s-})dA_1(s) \to \int_{[0, t]} f'(X_{s-})dA_1(s) \text{ a.s. as } n \to \infty \]
by the Dominated Convergence Theorem. Therefore, the first integral in the right-hand side of (A.3) indeed converges to the corresponding integral in (A.1). The convergence
\[ \sum_{0 < s \leq t} \{ f_n(X_s) - f_n(X_{s-}) - f_n'(X_{s-})\Delta X_s \} \to \sum_{0 < s \leq t} \{ f(X_s) - f(X_{s-}) - f'(X_{s-})\Delta X_s \}, \text{ } P_x\text{-a.s.} \]
follows from the discrete time version of the Dominated Convergence Theorem since \( |f_n(X_s) - f_n(X_{s-}) - f_n'(X_{s-})\Delta X_s| \) is bounded by \( \frac{C}{t}(\Delta X_s)^2 \) and \( \sum_{0 < s \leq t}(\Delta X_s)^2 < \infty \). Finally we need to show that
\[ \frac{\sigma^2}{2} \int_{[0, t]} f_n''(X_s)ds \to \frac{\sigma^2}{2} \int_{[0, t]} f''(X_s)ds, \text{ } P_x\text{-a.s.} \]
as \( n \to \infty \). If \( \sigma = 0 \) then there is nothing to prove and if \( \sigma \neq 0 \) then this follows from Lemma Appendix A.3 and the Dominated Convergence Theorem.

Appendix B. Appendix

In Appendix B we recall some results from [15] for the convenience of the reader. We minimally changed the notation in order to accommodate the present notations. It will be assumed in this section, just like in [15], that \( Y \) is a spectrally negative Lévy process, that is, \( \nu(0, \infty) = 0 \), and \( \sim Y \) is not a subordinator. In Section 4 we assumed that \( Y \) is spectrally positive, but it requires only minimal modification to adapt the results below to the spectrally positive case.

As explained at the beginning of Section 8.1 in [15], rather than working with the Lévy-Khintchine exponent, it is preferable to work with the Laplace exponent
\[ \psi(\lambda) := \frac{1}{t} \log E \left[ e^{\lambda Y(t)} \right], \]
which is finite for all \( \lambda \geq 0 \). The function \( \psi : [0, \infty) \to \mathbb{R} \) is zero at zero and tends to infinity at infinity. Further, it is infinitely differentiable and strictly convex. Define the right inverse
\[ \Phi(q) = \sup\{ \lambda \geq 0 : \psi(\lambda) = q \} \]
for each \( q \geq 0 \). If \( \psi'(0) \geq 0 \) then \( \lambda = 0 \) is the unique solution of \( \psi(\lambda) = 0 \) and otherwise there are two solutions to the latter with \( \lambda = \Phi(0) > 0 \) being the larger of the two, the other is \( \lambda = 0 \).

For all \( z \in \mathbb{R} \) let
\[ \tau^+_z = \inf\{ t > 0 : Y_t > z \} \text{ and } \tau^-_z = \inf\{ t > 0 : Y_t < z \}. \]
Next we recall Theorem 8.1 in [15].

Theorem Appendix B.1. (One- and two-sided exit formulae). There exists a family of functions \( W(q) : \mathbb{R} \to [0, \infty) \) and
\[ Z(q)(x) = 1 + q \int_0^x W(q)(y)dy, \text{ for } x \in \mathbb{R} \]
defined for each \( q \geq 0 \) such that the following hold.
(i) For any \( q \geq 0 \), we have \( W(q)(x) = 0 \) for \( x < 0 \) and \( W(q) \) is characterized on \([0, \infty)\) as a strictly increasing and continuous function whose Laplace transform satisfies
\[ \int_0^\infty e^{-\beta x} W(q)(x)dx = \frac{1}{\psi(\beta) - q} \text{ for } \beta > \Phi(q). \]
(ii) For any $x \in \mathbb{R}$ and $q \geq 0$

$$E_x \left[ \exp \left\{ -q \tau^+_a \right\} 1\{\tau^-_a < \infty\} \right] = Z(q)(x) - \frac{q}{\phi(q)}W(q)(x),$$

where we understand $q/\phi(q)$ in the limiting sense for $q = 0$, so that

$$P_x(\tau^-_a < \infty) = \begin{cases} 1 - \psi'(0+)W(0)(x) & \text{if } \psi'(0+) > 0 \\ 1 & \text{if } \psi'(0+) \leq 0. \end{cases}$$

(iii) For any $x \leq a$ and $q \geq 0$,

$$E_x \left[ \exp \{-q \tau^-_a \} 1\{\tau^-_a < \tau^+_a\} \right] = \frac{W(q)(x)}{W(q)(a)},$$

and

$$E_x \left[ \exp \{-q \tau^+_a \} 1\{\tau^-_a > \tau^+_a\} \right] = Z(q)(x) - \frac{Z(q)(a)}{W(q)(a)}W(q)(x).$$

Next we recall the definition of the $q$-potential measure of $Y$ killed on exiting $[0,a]$ as

$$U(q)(x,dy) = \int_0^\infty e^{-qt}P_x(Y_t \in dy, \tau > t)dt$$

for $q \geq 0$, where $a > 0$ and $\tau = \tau^+_a \wedge \tau^-_a$, with the agreement that $U(0) = 0$. We recall Theorem 8.7 in [15].

Theorem Appendix B.2. Suppose, for $q \geq 0$, that $U(q)(x,dy)$ is the $q$-potential measure of $Y$ killed on exiting $[0,a]$ where $x, y \in [0,a]$. Then it has a density $u(q)(x,y)$ given by

$$u(q)(x,y) = \frac{W(q)(x)W(q)(a-y)}{W(q)(a)} - W(q)(x-y).$$

References

[1] Søren Asmussen. *Applied Probability and Queues*, volume 2. Springer, 2003.

[2] Avner Bar-Ilan, David Perry, and Wolfgang Stadje. A generalized impulse control model of cash management. *Journal of Economic Dynamics and Control*, 28(6):1013–1033, 2004.

[3] Alain Bensoussan and Lions Jacques-Louis. Nouvelles formulation de problèmes de contrôle impulsionel et applications. *C.R. Acad. Sciences*, 276(A):1189–1192, 1973.

[4] Alain Bensoussan and Jacques-Louis Lions. Nouvelles méthodes en contrôle impulsionnel. *Applied Mathematics and Optimization*, 1(4):289–312, 1975.

[5] Alain Bensoussan, RH Liu, and Suresh P Sethi. Optimality of an (s,S) policy with compound Poisson and diffusion demands: A quasi-variational inequalities approach. *SIAM Journal on Control and Optimization*, 44(5):1650–1676, 2005.

[6] Abel Cadenillas, Peter Lakner, and Michael Pinedo. Optimal control of a mean-reverting inventory. *Operations Research*, 58(6):1697–1710, 2010.

[7] Abel Cadenillas, Sudipto Sarkar, and Fernando Zapatero. Optimal dividend policy with mean-reverting cash reservoir. *Mathematical Finance*, 17(1):81–109, 2007.

[8] George M Constantinides and Scott F Richard. Existence of optimal simple policies for discounted-cost inventory and cash management in continuous time. *Operations Research*, 26(4):620–636, 1978.

[9] Irmina Czarna and Adam Kaszubowski. Optimality of impulse control problem in refracted Lévy model with Parisian ruin and transaction costs. *Journal of Optimization Theory and Applications*, 185(3):982–1007, 2020.

[10] Haolin Feng and Kumar Muthuraman. A computational method for stochastic impulse control problems. *Mathematics of Operations Research*, 35(4):839–850, 2010.

[11] J Michael Harrison, Thomas M Sellke, and Allison J Taylor. Impulse control of Brownian motion. *Mathematics of Operations Research*, 8(3):454–466, 1983.

[12] Sheng-wo He, Jia-gang Wang, and Jia-an Yan. *Semimartingale Theory and Stochastic Calculus*. Science Press & CRC Press Inc., 1992.

[13] Jean Jacod and Albert Shiryaev. *Limit Theorems for Stochastic Processes*, volume 288. Springer Science & Business Media, 2013.
[14] Offer Kella and Ward Whitt. Useful martingales for stochastic storage processes with Lévy input. *Journal of Applied Probability*, 29(2):396–403, 1992.

[15] Andreas E Kyprianou. *Introductory Lectures on Fluctuations of Lévy Processes with Applications*. Springer Science & Business Media, 2006.

[16] Bernt Øksendal. *Stochastic Differential Equations: An Introduction with Applications*. Springer Science & Business Media, 2013.

[17] Melda Ormeci, J.G. Dai, and John Vande Vate. Impulse control of Brownian motion: The constrained average cost case. *Operations Research*, 56(3):618–629, 2008.

[18] Philip E Protter. *Stochastic Integration and Stochastic Differential Equations*. Springer, Berlin, 2004.

[19] Scott F Richard. Optimal impulse control of a diffusion process with both fixed and proportional costs of control. *SIAM Journal on Control and Optimization*, 15(1):79–91, 1977.

[20] Ken-Iti Sato. *Lévy Processes and Infinitely Divisible Distributions*. Cambridge University Press, 1999.

[21] Agnès Sulem. A solvable one-dimensional model of a diffusion inventory system. *Mathematics of Operations Research*, 11(1):125–133, 1986.

[22] Jinhiao Wu. Optimal exchange rates management using stochastic impulse control for geometric Lévy processes. *Mathematical Methods of Operations Research*, 89(2):257–280, 2019.

[23] Kazutoshi Yamazaki. Inventory control for spectrally positive Lévy demand processes. *Mathematics of Operations Research*, 42(1):212–237, 2017.