Parametric survival model to identify the predictors of breast cancer mortality: An accelerated failure time approach

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Background: Breast cancer (BC) was the fifth cause of mortality worldwide in 2015 and second cause of mortality in Iran in 2012. This study aimed to explore factors associated with survival of patients with BC using parametric survival models.

Materials and Methods: Data of 1154 patients that diagnosed with BC recorded in the East Azerbaijan population-based cancer registry database between March 2007 and March 2016. The parametric survival model with an accelerated failure time (AFT) approach was used to assess the association between sex, age, grade, and morphology with time to death. Results: A total of 217 (18.8%) individuals experienced death due to BC by the end of the study. Among the fitted parametric survival models including exponential, Weibull, log logistic, and log-normal models, the log-normal model was the best model with the Akaike information criterion = 1441.47 and Bayesian information criterion = 1486.93 where patients with higher ages (time ratio [TR] = 0.693; 95% confidence interval [CI] = [0.531, 0.904]) and higher grades (TR = 0.350; 95% CI = [0.201, 0.608]) had significantly lower survival while the lobular carcinoma type of morphology (TR = 1.975; 95% CI = [1.049, 3.720]) had significantly higher survival. Conclusion: Log-normal model showed to be an optimal tool to model the survival of patients with BC in the current study. Age, grade, and morphology showed significant association with time to death in patients with BC using AFT model. This finding could be recommended for planning and health policymaking in patients with BC. However, the impact of the models used for analysis on the significance and magnitude of estimated effects should be acknowledged.

Key words: Accelerated failure time, breast cancer, parametric model, survival

INTRODUCTION

Breast cancer (BC) is one of the prevalent cancers that is created by growth and spread of malignant cells from the breast tissue.¹ It is a common cancer among women and the most common reason of cancer-related mortality.² BC has the highest prevalence in the USA and Western Europe and the lowest prevalence in East Asia.² Despite significant reductions in BC prevalence and mortality in many countries, its global prevalence is increasing, and its age-standardized incidence rate increase 12% annually.²⁻⁵

The prevalence and mortality of BC are increasing in Iran and East Azerbaijan province in the northwest of Iran as well. Even though the BC incidence rates in East Azerbaijan are lower than the USA and Western Europe, the increase in BC incidence rates urges further study.⁶⁻⁷

BC has complex etiology with different risk factors such as heredity, hormonal factors, environmental factors, nutrition, number of pregnancy, age, tumor grade, and morphology.⁸⁻¹⁰ Assessing the risk factors and proper planning are important to reduce BC incidence and mortality.

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Prediction of patients’ survival and identifying high-risk patients are important for health management.\cite{11} Parametric survival models, among the various predictive models, use data distribution information to smooth the fluctuation caused by sampling error and may provide better estimates than those obtained by Cox model. Some of the parametric models use accelerated failure time (AFT) approach that directly targets the patients’ survival probabilities as an important clinical quantity. These models do not require proportional hazards (PH) assumption which may not be satisfied in most practical settings.\cite{12} Hence, using AFT models could provide better estimates and fit in some situations. More optimal and exact estimates could provide a better measure for decision-making by the users of the results of this study.

Based on an extensive search in the literature, there was no study, if any, to investigate the predictors of survival in patients with BC, so in this study, we used some parametric survival models with AFT approach to investigate the relationship between the survival of patients with BC and common prognostic factors.

**MATERIALS AND METHODS**

**Study population**

The data of patients diagnosed with primary BC were used, which have been collected from March 20, 2007, to March 19, 2016, by the population-based cancer registry of East Azerbaijan province of Iran. The BCs have been confirmed according to C50.0–9 codes of the International Classification of Disease-Oncology for the patients.\cite{13}

**Data collection**

In the East Azerbaijan population-based cancer registry (EA-PBCR), all clinic-pathologic data were registered regularly including age, sex, morphology (including ductal carcinoma [DC]; lobular carcinoma [LC]; and other types), and Grade (I–IV) using 11-digit personal identification number. Information about survival and outcome was obtained by contacting patients or their relatives or referring to the hospital information system during follow-up.

**Statistical analysis**

Data were expressed as frequency (percentages) for categorical and mean ± standard deviation for numeric variables. Survival, the primary outcome of our study, was calculated as the time from diagnosis of BC to death from BC or being censored by the end of the study period.

The PH assumption was checked for variables, and since it was not satisfied for grade, the parametric models were applied. We fitted parametric models including exponential (survival function defined as $S(t) = \exp[-\lambda t]$), Weibull ($S(t) = \exp[-\lambda t^p]$), log logistic ($S(t) = \frac{1}{1 + \lambda t^p}$), and log-normal ($S(t)=1-\phi \left( \frac{\ln t - \mu}{\sigma} \right)$). In AFT models, time ratio (TR) (sometimes called acceleration factor) is usually reported to reflect the impact of each variable on the survival. If TR for a variable is $\alpha$, it means that the lifespan of this group, on average, is stretched out $\alpha$-times longer than the lifespan of the reference group. AFT models assumed that the hazard of an event is constant over time. The assumption was assessed by the graphs of hazards over time. Having a constant increase in hazard overtime for all predictors confirmed this assumption.

All the variables were entered simultaneously in the multivariable model to obtain adjusted effects of the factors on the survival of patients. Statistical significance was set at 0.05. The fit of models was compared using probability plots and Akaike information criterion (AIC) and Bayesian information criterion (BIC), smaller value of which indicate better fit.

In addition, a sensitivity analysis was conducted to assess the effect of low number on men’s in the study sample, so that the above mentioned optimal model was performed just on the women’s data.

All analyses were performed using STATA MP 15.0 (StataCorp. 2017. Stata Statistical Software: Release 15. College Station, TX: Stata Corp LLC).

**RESULTS**

**Participant**

A total of 1132 patients were female (98.1%) and mean age of patients was 50.4 (SD 12.5) years. The median (25th, 75th percentiles) of survival time was 46.83 (32.57, 69.28) months. A total of 217 (18.8%) individuals experienced death due to BC by the end of the study. The 1-, 3-, and 5-year survival rate were 96.1% (95% confidence interval [CI]: 94.81–97.07), 88.2% (95% CI: 86.10–89.97), and 81.4% (95% CI: 78.55–83.85), respectively.

**Comparing mortality rate**

The mortality rate from BC was 18.6% in females and 27.3% in males ($P$ for difference = 0.305). The 1-, 3-, and 5-year survival rate was 96.2% (95% CI: 94.91–97.17), 88.2% (95% CI: 86.13–90.03), and 81.4% (95% CI: 78.58–83.93), respectively, in females and 91.9% (95% CI: 68.30–97.65), 86.4% (95% CI: 63.44–95.39), and 79.2% (95% CI: 52.35–91.91), respectively, in males.

Total percentage of BC death was 16.5 within patients aged < 50 years and 21.2 within patients aged ≥50 years. The 1-, 3-, and 5-year survival rate was 97.6%
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(95% CI: 95.98–98.57), 90.1% (95% CI: 87.24–92.33), and 84.4% (95% CI: 80.62–87.44), respectively, for patients aged <50 years and 94.6% (95% CI: 92.37–96.15), 86.2% (95% CI: 83.03–88.88), and 78.2% (95% CI: 73.84–82.00), respectively, for patients aged ≥50 years.

Furthermore, the mortality rate from BC was 2.1 within patients in Grade I, 14.7 within patients in Grade II, 26.9 within patients in Grade III, and 28.2 within patients in Grade IV. Furthermore, this rate was 17.8 within patients with DC type of morphology, 14.1 within patients with LC type and 31.1 within patients with other types.

Assessing the proportional hazards assumption
The PH assumption was checked for the variables, and it was confirmed for sex ($P = 0.35$), age ($P = 0.31$) and morphology ($P = 0.12$ and 0.49, respectively, for DC and other types of morphology), and was not confirmed for grade ($P = 0.36, 0.04$, and 0.02, respectively, for Category II, III, and IV).

Comparing the parametric models
According to AIC and BIC values [Table 1] and probability plots for survival data [Figure 1], the survival model with log-normal distribution had the best fit on the data.

The cumulative hazards of patients with BC are plotted against survival time in Figure 2.

Figure 2a indicates that the cumulative hazard for males is higher than females, but the difference was not significant ($P = 0.505$). As shown in Figure 2b, the cumulative hazard for patients aged ≥50 years is higher than that one for patients aged <50 years, and the difference was significant ($P = 0.021$). The differences in cumulative hazard among the morphology type groups were significant ($P < 0.001$). Figure 2c shows that the cumulative hazard for patients in LC type of morphology is lower than DC type, and the patients in other types of morphology had the highest cumulative hazard. The difference in cumulative hazard among the tumor grade groups was significant ($P < 0.001$), where it was the lowest for Grade I and the highest for Grade III and IV [Figure 2d] and the difference between Grade III and Grade IV was not significant ($P = 0.464$).

According to the results of parametric regressions, higher age (≥50), higher grades and DC, and other morphologies (compared to LC) had significantly lower survival in BC patients [Table 2].

The results of sensitivity analysis to consider the effect of low number on men’s in the study sample were presented in Table 3. The results did not differ to those of analyses by total sample, and approximately the same values of TRs have been observed for risk factors.

DISCUSSION

In this study, we investigated the underlying factors of survival in patients diagnosed with primary BC through

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Table 1: Akaike information criterion and Bayesian information criterion values for parametric models

| Distribution   | AIC       | BIC       |
|----------------|-----------|-----------|
| Exponential    | 1460.79   | 1501.20   |
| Weibull        | 1460.05   | 1505.52   |
| Log logistic   | 1454.37   | 1499.83   |
| Log-normal     | 1441.47   | 1486.93   |

AIC=Akaike information criterion; BIC=Bayesian information criterion

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Figure 1: Probability plots for breast cancer data

Figure 2: Cumulative hazards of patients with BC

Figure 2a: Males vs. females

Figure 2b: Patients aged ≥50 vs. <50 years

Figure 2c: LC vs. DC morphology

Figure 2d: Grade III vs. Grade IV
a parametric survival modeling approach. Interestingly, in our data, log-normal model had the best performance among all investigated parametric models. In line with our finding, log-normal model has outperformed other parametric models in modeling of progression-free survival in patients with advanced BC.[14,15] While Baghestani et al. concluded that Weibull model is a better model in modeling the survival in patients with colorectal cancer.[16] Parametric survival models may offer optimal models due to using information of data distribution. Furthermore, parametric survival models provide smooth graphs for survival and hazard functions and hence better estimation of such quantities specially in small sample studies or in the case of times with sparse data.[12,17] Although our data were right-censored, parametric survival models can easily accommodate left- and interval-censored data as well.[12] Furthermore, the PH assumption is not required to be assessed in some parametric model such as log-normal, log-logistic and generalized gamma do not depend on the PH assumption. In addition, in parametric models, the parameters can be estimated and completely

| Variable | Exponential | | | Weibull | | | | Log logistic | | | Log normal | | |
|----------|-------------|------------|----------|------------|----------|----------|------------|-------------|----------|----------|
|          | Time ratio (95% CI) | P | Time ratio (95% CI) | P | Time ratio (95% CI) | P | Time ratio (95% CI) | P |
| Age      |             |          |          |             |          |          |             |          |
| <50      | Reference   |          | Reference |          |          |          | Reference   |          |
| ≥50      | 0.72 (0.55-0.95) | 0.019* | 0.74 (0.58-0.95) | 0.018* | 0.72 (0.56-0.93) | 0.012* | 0.69 (0.53-0.90) | 0.007* |
| Sex      |             |          |          |             |          |          |             |          |
| Female   | Reference   |          | Reference |          |          |          | Reference   |          |
| Male     | 0.81 (0.36-1.84) | 0.617 | 0.83 (0.40-1.75) | 0.635 | 0.80 (0.35-1.82) | 0.590 | 0.71 (0.30-1.69) | 0.439 |
| Grade    |             |          |          |             |          |          |             |          |
| I        | Reference   |          | Reference |          |          |          | Reference   |          |
| II       | 0.65 (0.42-1.01) | 0.057 | 0.67 (0.45-0.99) | 0.047* | 0.65 (0.44-0.97) | 0.035* | 0.64 (0.43-0.96) | 0.031* |
| III      | 0.49 (0.31-0.76) | 0.002* | 0.53 (0.35-0.80) | 0.003* | 0.49 (0.32-0.74) | 0.001* | 0.44 (0.28-0.67) | <0.001* |
| IV       | 0.38 (0.22-0.66) | 0.001* | 0.42 (0.25-0.69) | 0.001* | 0.38 (0.23-0.66) | <0.001* | 0.35 (0.20-0.61) | <0.001* |
| Morphology |             |          |          |             |          |          |             |          |
| DC       | Reference   |          | Reference |          |          |          | Reference   |          |
| LC       | 1.72 (0.89-3.34) | 0.106 | 1.63 (0.89-2.97) | 0.114 | 1.76 (0.96-3.25) | 0.068 | 1.97 (1.05-3.72) | 0.035* |
| Other    | 0.65 (0.43-0.98) | 0.039* | 0.66 (0.45-0.96) | 0.029* | 0.65 (0.43-0.97) | 0.038* | 0.61 (0.40-0.95) | 0.030* |

*P<0.05. CI=Confidence interval; DC=Ductal carcinoma; LC=Lobular carcinoma

Figure 2: Cumulative hazard of patients with breast cancer by sex (a), no significant difference between males and females: P =0.505), by age (b), significant difference between higher and lower age: P =0.021), by morphology (c), significant difference among morphology types: P <0.001) and by grade (d), significant difference among different grades: P <0.001)
specify the survival and hazard functions. This simplicity and completeness are the other main appeals of using a parametric approach.\cite{12}

Like our study, Foerster \textit{et al.} found that sex had no significant relationship with survival.\cite{13}

Moreover, the age was inversely related to the survival of patients with BC. Previous studies have similarly revealed that the survival in patients with BC was correlated with their age.\cite{19} Correspondingly another study indicated that the risk of mortality had been increased by patients’ age.\cite{9,20,21} Old patients may have lower survival than younger patients due to the existence of other possible persistent diseases, inappropriate health status, and diagnosis in a higher stage.\cite{20} On the other hand, some other studies did not indicate any significant relationship between patients’ survival and their age.\cite{22-24}

The grade was another significant-related factor with survival so that an inverse relationship was observed between this factor and survival of patients. Similarly, previous studies have also discovered that patients’ survival was inversely connected by the grade.\cite{9,19,22,25} It was concluded that, by increasing the tumor grade, the tumor becomes malignant, and therefore the patients’ survival decreases.

Morphology was another significantly related factor of survival in patients with BC. The LC and DC types of morphology do not have any significant difference. The LC type had the highest survival compared to other types of morphology. Similar to our findings, Møller \textit{et al.} found that LC type of morphology and Fallahzadeh \textit{et al.} found that the DC and LC types of morphology had significantly higher survival in patients with BC.\cite{9,20}

\section*{Strengths and limitations}

Obviously, our study was not without limitations. Some important clinical information such as tumor-node-metastasis stage, tumor size, estrogen receptor, and progesterone receptor were not available to be include in the analysis and improve the fit of models. As another limitation, in Iran, the Iranian National Cancer Registry and EA-PBCR have not yet developed. In the EA-PBCR, to avoid missing data, it was tried to collect data through a combined active and passive follow-up protocol and register all patients with newly diagnosed cancer from across the province. Because of the lack of contact information, the data for some of the initial cases were missed. This study was limited to recent 10 years because of the lag time in data reporting and registration.

\section*{CONCLUSIONS}

Interestingly, using the parametric survival model, we achieved more flexibility to model the predictors of survival in patients with BC. Using log-normal parametric survival model led to remarkable advantageous AFT parameterization which could be interpreted based on survival (not hazard). The model showed that age, grade, and morphology were significant predictors of survival. These findings could be recommended for prevention, planning, health policymaking, early diagnosis of BC and early treatment and so increase survival in patients with BC.

\section*{Acknowledgments}

Data of patients diagnosed with primary Breast cancer and registered in the East Azerbaijan population-based cancer registry were included in our analysis. As the ethics rules of East Azerbaijan population-based cancer registry, all patients’ information, and records are confidential. The study protocol was approved by the Institutional Review Board of Tabriz University of Medical Sciences (IRB no.: IR.TBZMED.REC.1397.986).

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\section*{Conflicts of interest}

There are no conflicts of interest.

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