Communication-Efficient Triangle Counting under Local Differential Privacy

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Abstract

Triangle counting in networks under LDP (Local Differential Privacy) is a fundamental task for analyzing connection patterns or calculating a clustering coefficient while strongly protecting sensitive friendships from a central server. In particular, a recent study proposes an algorithm for this task that uses two rounds of interaction between users and the server to significantly reduce estimation error. However, this algorithm suffers from a prohibitively high communication cost due to a large noisy graph each user needs to download.

In this work, we propose triangle counting algorithms under LDP with a small estimation error and communication cost. We first propose two-rounds algorithms consisting of edge sampling and carefully selecting edges each user downloads so that the estimation error is small. Then we propose a double clipping technique, which clips the number of edges and then the number of noisy triangles, to significantly reduce the sensitivity of each user’s query. Through comprehensive evaluation, we show that our algorithms dramatically reduce the communication cost of the existing algorithm, e.g., from 6 hours to 8 seconds or less at a 20 Mbps download rate, while keeping a small estimation error.

1 Introduction

Counting subgraphs (e.g., triangles, stars, cycles) is one of the most basic tasks for analyzing connection patterns in various graph data, e.g., social, communication, and collaboration networks. For example, a triangle is given by a set of three nodes with three edges, whereas a k-star is given by a central node connected to k other nodes. These subgraphs play a crucial role in calculating a clustering coefficient (see Figure 1). The clustering coefficient measures the average probability that two friends of a user will also be friends in a social graph [45]. Therefore, it is useful for measuring the effectiveness of friend suggestions. In addition, the clustering coefficient represents the degree to which users tend to cluster together. Thus, if it is large in some services/communities, we can effectively apply social recommendations [38] to the users. Triangles and k-stars are also useful for constructing graph models [31, 52]; see also [59] for other applications of triangle counting. However, graph data often involve sensitive data such as sensitive edges (friendships), and they can be leaked from exact numbers of triangles and k-stars [30].

To analyze subgraphs while protecting user privacy, DP (Differential Privacy) [25] has been widely adopted as a privacy metric [24, 30, 36, 64, 65, 67]. DP protects user privacy against adversaries with arbitrary background knowledge and is known as a gold standard for data privacy. According to the underlying model, DP can be categorized into central (or global) DP and LDP (Local DP). Central DP assumes a scenario where a central server has personal data of all users. Although accurate analysis of subgraphs is possible under this model [24, 36, 67], there is a risk that the entire graph is leaked from the server by illegal access or internal fraud [18, 43]. In addition, central DP cannot be applied to decentralized social networks [4–6, 48] where the entire graph is distributed across many servers. We can even consider fully decentralized applications where a server does not have any original edge, e.g., a mobile app that sends a noisy degree (noisy number of friends) to the server, which then estimates a degree distribution. Central DP cannot be used in such applications.

In contrast, LDP assumes a scenario where each user obfuscates her personal data (friends list in the case of graphs) by herself and sends the obfuscated data to a possibly malicious server; i.e., it does not assume trusted servers. Thus, it does not suffer from a data breach and can also be applied to the decentralized applications. LDP has been widely studied in tabular data where each row corresponds to a user’s personal data (e.g., age, browser setting, location) [8, 12, 27, 33, 44, 60] and also in graph data [30, 49, 64, 65]. For example, k-star counts can be very accurately estimated under LDP because each user can count k-stars of which she is a center and sends a noisy version of her k-star count to the server [30].

However, more complex subgraphs such as triangles are much harder to count under LDP because each user cannot see

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Our Contributions. We propose locally private triangle counting algorithms with a small estimation error and small communication cost. Our contributions are as follows:

- We propose two-rounds triangle algorithms consisting of edge sampling after RR and selecting edges each user downloads. In particular, we show that a simple extension of [30] with edge sampling suffers from a large estimation error for a large or dense graph where the number of 4-cycles (such as \(v_1-v_2-v_3-v_6-v_1\) in Figure 1) is large. To address this issue, we propose some strategies for selecting edges to download to reduce the error caused by the 4-cycles, which we call the 4-cycle trick.

- We show that the algorithms with the 4-cycle trick still suffer from a large estimation error due to large Laplacian noise for each user. To significantly reduce the Laplacian noise, we propose a double clipping technique, which clips a degree (the number of edges) of each user with LDP and then clips the number of noisy triangles.

- We evaluate our algorithms using two real datasets. We show that our entire algorithms with the 4-cycle trick and double clipping dramatically reduce the communication cost of [30]. For example, for a graph with about 900000 users, we reduce the download cost from 400 Gbits (6 hours when 20 Mbps) to 160 Mbits (8 seconds) or less while keeping the relative error much smaller than 1.

Thus, locally private triangle counting is now much more practical. In Appendix C, we also show that we can estimate the clustering coefficient with a small estimation error and download cost. For example, our algorithms are useful for measuring the effectiveness of friend suggestions or social recommendations in decentralized social networks, e.g., Diaspora [4], Mastodon [5]. Our source code is available at [1].

All the proofs of our privacy and utility analysis are given in Appendices F, G, and H.

Technical Novelty. Below we explain more about the technical novelty of this paper. Although we focus on two-rounds local algorithms in the same way as [30], we introduce several new algorithmic ideas previously unknown in the literature.

First, our 4-cycle trick is totally new. Although some studies focus on 4-cycle counting [13, 35, 40, 42], this work is the first to use 4-cycles to improve communication efficiency. Second, selective download of parts of a centrally computed quantity is also new. This is not limited to graphs – even in machine learning, there are no such strategic download techniques previously, to our knowledge. Third, our utility analysis of our triangle algorithms (Theorem 2) is different from [30] in that ours introduces subgraphs such as 4-cycles and k-stars. This leads us to our 4-cycle trick. Fourth, we propose two triangle algorithms that introduce the 4-cycle trick and show that the more tricky one provides the best performance because of its low sensitivity in DP.

Finally, our double clipping is new. Andrew et al. [9] propose an adaptive clipping technique, which applies clipping twice. However, they focus on federated averaging, and their problem setting is different from our graph setting. In particular, they require a private quantile of the norm distribution. In contrast, we need only a much simpler estimate: a private degree. Here, we use the fact that the degree has a small sensitivity (sensitivity = 1) in DP for edges. We also provide a new, reasonably tight bound on the probability that the noisy triangle count exceeds a clipping threshold (Theorem 4). Thanks to the two differences, we obtain a significant communication improvement: two or three orders of magnitude.
2 Related Work

Triangle Counting. Triangle counting has been extensively studied in a non-private setting [14,15,21,26,54,57,58,62] (it is almost a sub-field in itself) because it requires high time complexity for large graphs.

Edge sampling [14,26,58,62] is one of the most basic techniques to improve scalability. Although edge sampling is simple, it is quite effective — it is reported in [62] that edge sampling outperforms other sampling techniques such as node sampling and triangle sampling. Based on this, we adopt edge sampling after RR\(^1\) with new techniques such as the 4-cycle trick and double clipping. Our entire algorithms significantly improve the communication cost, as well as the space and time complexity, under LDP (see Sections 5.3 and 6).

DP on Graphs. For private graph analysis, DP has been widely adopted as a privacy metric. Most of them adopt central (or global) DP [23,24,29,36,37,50,67], which suffers from the data breach issue.

LDP on graphs has recently studied in some studies, e.g., synthetic data generation [49], subgraph counting [30,56,64,65]. A study in [56] proposes subgraph counting algorithms in a setting where each user allows her friends to see all her connections. However, this setting is unsuitable for many applications; e.g., in Facebook, a user can easily change her setting so that her friends cannot see her connections.

Thus, we consider a model where each user can see only her friends. In this model, some one-round algorithms [64,65] and two-rounds algorithms [30] have been proposed. However, they suffer from a prohibitively large estimation error or high communication cost, as explained in Section 1.

Recently proposed network LDP protocols [22] consider, instead of a central server, collecting private data with user-to-user communication protocols along a graph. They focus on sums, histograms, and SGD (Stochastic Gradient Descent) to improve scalability. Although edge sampling is not strictly edge sampling, our methods can be applied to our task of subgraph counting under LDP for edges. The same applies to another work [53] that improves the utility of an averaging query by correlating the noise of users according to a graph.

LDP. RR [33,61] and RAPPOR [27] have been widely used for tabular data in LDP. Our work uses RR in part of our algorithm but builds off of it significantly. One noteworthy result in this area is HR (Hadamard Response) [8], which is state-of-the-art for tabular data and requires low communication. However, this result is not applied to graph data and does not address the communication issues considered in this paper. Specifically, applying HR to each bit in a neighbor list will result in \(O(n^2)\) (\(n\) : #users) download cost in the same way as the previous work [30] that uses RR. Applying HR to an entire neighbor list (which has \(2^n\) possible values) will similarly result in \(O(n \log 2^n) = O(n^2)\) download cost.

Previous work on distribution estimation [33,44,60] or heavy hitters [12] addresses a different problem than ours, as they assume that every user has i.i.d. (independent and identically distributed) samples. In our setting, a user’s neighbor list is non-i.i.d. (as one edge is shared by two users), which does not fit into their statistical framework.

3 Preliminaries

3.1 Notations

We begin with basic notations. Let \(\mathbb{N}, \mathbb{R}, \mathbb{Z}_{\geq 0},\) and \(\mathbb{R}_{\geq 0}\) be the sets of natural numbers, real numbers, non-negative integers, and non-negative real numbers, respectively. For \(z \in \mathbb{N},\) let \([z]\) be a set of natural numbers from 1 to \(z; i.e.,[z] = \{1, 2, ..., z\}\).

Let \(G = (V,E)\) be an undirected graph, where \(V\) is a set of nodes and \(E \subseteq V \times V\) is a set of edges. Let \(n \in \mathbb{N}\) be the number of nodes in \(V\). Let \(v_i \in V\) be the \(i\)-th node; i.e., \(V = \{v_1, ..., v_n\}\). We consider a social graph where each node in \(V\) represents a user and an edge \((v_i,v_j) \in E\) represents that \(v_i\) is a friend with \(v_j\). Let \(d_{max} \in \mathbb{N}\) be the maximum degree of \(G\). Let \(G\) be a set of graphs with \(n\) nodes. Let \(f_{\Delta} : G \rightarrow \mathbb{Z}_{\geq 0}\) be a triangle count query that takes \(G \in G\) as input and outputs a triangle count \(f_{\Delta}(G)\) (i.e., number of triangles in \(G\)).

Let \(A = (a_{i,j}) \in \{0,1\}^{n \times n}\) be a symmetric adjacency matrix corresponding to \(G; i.e., a_{i,j} = 1\) if and only if \((v_i,v_j) \in E\). We consider a local privacy model [30,49], where each user obfuscates her neighbor list \(a_i = (a_{i,1}, ..., a_{i,n}) \in \{0,1\}^n\) (i.e., the \(i\)-th row of \(A\)) using a local randomizer \(R_v\) with domain \(\{0,1\}^n\) and sends obfuscated data \(R_v(a_i)\) to a server. We also assume a two-rounds algorithm in which user \(v_i\) downloads a message \(M_i\) from the server at the second round.

We also show the basic notations in Table 2 of Appendix A.

3.2 Local Differential Privacy on Graphs

LDP on Graphs. When we apply LDP (Local DP) to graphs, we follow the direction of edge DP [47,51] that has been developed for the central DP model. In edge DP, the existence of an edge between any two users is protected; i.e., two computations, one using a graph with the edge and one using the graph without the edge, are indistinguishable. There is also another privacy notion called node DP [29,66], which hides the existence of one user along with all her edges. However, in the local model, many applications send a user ID to a server; e.g., each user sends the number of her friends along with her user ID. For such applications, we cannot use node DP but can use edge DP to hide her edges, i.e., friends. Thus,
we focus on edge DP in the local model in the same way as [30, 49, 56, 64, 65].

Specifically, assume that user \( v_i \) uses her local randomizer \( R_i \). We assume that the server and other users can be honest-but-curious adversaries and that they can obtain all edges except for user \( v_i \)'s edges as prior knowledge. Then we use the following definition for \( R_i \):

**Definition 1** (\( \epsilon \)-edge LDP [49]). Let \( \epsilon \in \mathbb{R}_{\geq 0} \). For \( i \in [n] \), let \( R_i \) be a local randomizer of user \( v_i \) that takes \( a_i \) as input. We say \( R_i \) provides \( \epsilon \)-edge LDP if for any two neighbor lists \( a_i, a'_i \in \{0,1\}^s \) that differ in one bit and any \( s \in \text{Range}(R_i) \),

\[
\Pr[R_i(a_i) = s] = e^\epsilon \Pr[R_i(a'_i) = s].
\]

(1)

For example, a local randomizer \( R_i \) that applies Warner’s RR (Randomized Response) [61], which flips 0/1 with probability \( \frac{1}{2} - \epsilon \), to each bit of \( a_i \) provides \( \epsilon \)-edge LDP.

The parameter \( \epsilon \) is called the privacy budget. When \( \epsilon \) is small (e.g., \( \epsilon \leq 1 \) [39]), each bit is strongly protected by edge LDP. Edge LDP can also be used to hide multiple bits – by group privacy [25], two neighbor lists \( a_i, a'_i \in \{0,1\}^n \) that differ in \( k \in \mathbb{N} \) bits are indistinguishable up to the factor \( e^k \).

Edge LDP is useful for protecting a neighbor list \( a_i \) of each user \( v_i \). For example, a user in Facebook can change her setting so that anyone (except for the central server) cannot see her friend list \( a_i \). Edge LDP hides \( a_i \) even from the server.

As with regular LDP, the guarantee of edge LDP does not break even if the server or other users act maliciously. However, adding or removing an edge affects the neighbor list of two users. This means that each user needs to trust her friend to not reveal an edge between them. This also applies to Facebook – even if \( v_i \) keeps \( a_i \) secret, her edge with \( v_j \) can be disclosed if \( v_j \) reveals \( a_j \). To protect each edge during the whole process, we use another privacy notion called relationship DP [30]:

**Definition 2** (\( \epsilon \)-relationship DP [30]). Let \( \epsilon \in \mathbb{R}_{\geq 0} \). For \( i \in [n] \), let \( R_i \) be a local randomizer of user \( v_i \) that takes \( a_i \) as input. We say \( (R_1, \ldots, R_n) \) provides \( \epsilon \)-relationship DP if for any two neighboring graphs \( G, G' \in \mathcal{G} \) that differ in one edge and any \( (s_1, \ldots, s_n) \in \text{Range}(R_1) \times \ldots \times \text{Range}(R_n) \),

\[
\Pr[(R_1(a_1), \ldots, R_n(a_n)) = (s_1, \ldots, s_n)] \leq e^\epsilon \Pr[(R_1(a'_1), \ldots, R_n(a'_n)) = (s_1, \ldots, s_n)],
\]

(2)

where \( a_i \) (resp. \( a'_i \)) \in \{0,1\}^n \) is the \( i \)-th row of the adjacency matrix of graph \( G \) (resp. \( G' \)).

If users \( v_i \) and \( v_j \) follow the protocol, (2) holds for graphs \( G, G' \) that differ in \( (v_i, v_j) \). Thus, relationship DP applies to all edges of a user whose neighbors are trustworthy.

While users need to trust other friends to maintain a relationship DP guarantee, only one edge per user is at risk for each malicious friend that does not follow the protocol. This is because only one edge can exist between two users. Thus, although the trust assumption in relationship DP is stronger than that of LDP, it is much weaker than that of central DP in which all edges can be revealed by the server.

It is possible to use a tuple of local randomizers with edge LDP to obtain a relationship DP guarantee:

**Proposition 1** (Edge LDP and relationship DP [30]). If each of local randomizers \( R_1, \ldots, R_n \) provides \( \epsilon \)-edge LDP, then \( (R_1, \ldots, R_n) \) provides \( 2\epsilon \)-relationship DP. Alternatively, if each \( R_i \) uses only bits \( a_{i1}, \ldots, a_{i,k-1} \) for users with smaller IDs (i.e., only the lower triangular part of \( A \)), then \( (R_1, \ldots, R_n) \) provides \( \epsilon \)-relationship DP.

The doubling factor in \( \epsilon \) comes from the fact that (2) applies to an entire edge, whereas (1) applies to just one neighbor list, and adding an entire edge may cause changes to two neighbor lists. However, if each \( R_i \) ignores bits \( a_{i,k}, \ldots, a_{i,n} \) for users with larger IDs, then this doubling factor can be avoided. Our algorithms also use only the lower triangular part of \( A \) to avoid this doubling issue.

**Interaction among Users and Multiple Rounds.** While interaction in LDP has been studied before [32], neither of Definitions 1 and 2 allows the interaction among users in a one-round protocol where user \( v_i \) sends \( R_i(a_i) \) to the server.

However, the interaction among users is possible in a multi-round protocol. Specifically, at the first round, user \( v_i \) applies a randomizer \( R_i \) and sends \( R(a_i) \) to the server. At the second round, the server calculates a message \( M_i \) for \( v_i \) by performing some post-processing on \( R(a_i) \), possibly with the private outputs by other users. Let \( \lambda_i \) be the post-processing algorithm on \( R(a_i) \); i.e., \( M_i = \lambda_i(R(a_i)) \). The server sends \( M_i \) to \( v_i \). Then, \( v_i \) uses a randomizer \( R(M_i) \) that depends on \( M_i \) and sends \( R(M_i)(a_i) \) back to the server. This entire computation provides DP by a (general) sequential composition [39]:

**Proposition 2** (Sequential composition of edge LDP). For \( i \in [n] \), let \( M_i \) be a local randomizer of user \( v_i \) that takes \( a_i \) as input. Let \( \lambda_i \) be a post-processing algorithm on \( R(a_i) \), and \( M_i = \lambda_i(R(a_i)) \) be its output. Let \( R(M_i) \) be a local randomizer of \( v_i \) that depends on \( M_i \). If \( R(a_i) \) provides \( \epsilon_1 \)-edge LDP and for any \( M \in \text{Range}(\lambda_i) \), \( R(M_i)(a_i) \) provides \( \epsilon_2 \)-edge LDP, then the sequential composition \( (R(a_i), R(M_i)(a_i)) \) provides \( (\epsilon_1 + \epsilon_2) \)-edge LDP.

We provide a proof of Proposition 2 in Appendix F.

**Global Sensitivity.** We use the notion of global sensitivity [25] to provide edge LDP:

**Definition 3.** In edge LDP (Definition 1), the global sensitivity of a function \( f : \{0,1\}^n \rightarrow \mathbb{R} \) is given by:

\[
GS_f = \max_{a_i, a'_i \in \{0,1\}, a_i \sim a'_i} |f(a_i) - f(a'_i)|,
\]

where \( a_i \sim a'_i \) represents that \( a_i \) and \( a'_i \) differ in one bit.

For example, adding the Laplacian noise with mean 0 scale \( \frac{GS_f}{\epsilon} \) (denoted by \( \text{Lap}(\frac{GS_f}{\epsilon}) \)) to \( f(a_i) \) provides \( \epsilon \)-edge LDP.
3.3 Utility and Communication-Efficiency

Utility. We consider a private estimate of \( f_\Delta(G) \). Our private estimator \( \hat{f}_\Delta : G \to \mathbb{R} \) is a post-processing of local randomizers \((R_1, \ldots, R_n)\) that satisfy \(\varepsilon\)-edge LDP. Following previous work, we use the \( l_2 \) loss (i.e., squared error) \([33, 44, 60]\) and the relative error \([16, 19, 63]\) as utility metrics.

Specifically, let \( l_2^2 \) be the expected \( l_2 \) loss function on a graph \( G \), which maps the estimate \( \hat{f}_\Delta(G) \) and the true value \( f_\Delta(G) \) to the expected \( l_2 \) loss; i.e., \( l_2^2(f_\Delta(G), \hat{f}_\Delta(G)) = \mathbb{E}[(\hat{f}_\Delta(G) - f_\Delta(G))^2] \). The expectation is taken over the randomness in the estimator \( \hat{f} \), which is necessarily a randomized algorithm since it satisfies edge LDP. In our theoretical analysis, we analyze the expected \( l_2 \) loss, as with \([33, 44, 60]\).

Communication-Efficiency. A prominent concern when performing local computations is that the computing power of individual users is often limited. Of particular concern to our private estimators, and a bottleneck of previous work in locally private triangle counting \([30]\), is the communication overhead between users and the server. This communication takes the form of users downloading any necessary data required to compute their local randomizers and uploading the output of their local randomizers. We distinguish the two quantities because often downloading is cheaper than uploading.

Consider a \( \tau \)-round protocol, where \( \tau \in \mathbb{N} \). At round \( j \in [\tau] \), user \( v_i \) applies a local randomizer \( R_i^j(M_i^j) \) to her neighbor list \( a_i \), where \( M_i^j \) is a message sent from the server to user \( v_i \) during round \( j \). We define the download cost as the number of bits required to describe \( M_i^j \) and the upload cost as the number of bits required to describe \( R_i^j(M_i^j)(a_i) \). Over all rounds and all users, we evaluate the maximum per-user download/upload cost, which is given by:

\[
\text{Cost}_{DL} = \max_{i=1}^{n} \sum_{j=1}^{\tau} \mathbb{E}[|M_i^j|] \quad \text{(bits)} \quad (3)
\]

\[
\text{Cost}_{UL} = \max_{i=1}^{n} \sum_{j=1}^{\tau} \mathbb{E}[|R_i^j(M_i^j)(a_i)|] \quad \text{(bits)} \quad (4)
\]

The above expectations go over the probability distributions of computing the local randomizers and any post-processing done by the server. We evaluate the maximum of the expected download/upload cost over users.

4 Communication-Efficient Triangle Counting Algorithms

The current state-of-the-art triangle counting algorithm \([30]\) under edge LDP suffers from an extremely large per-user download cost; e.g., every user has to download a message of 400 Gbits or more when \( n = 900000 \). Therefore, it is impractical for a large graph. To address this issue, we propose three communication-efficient triangle algorithms under edge LDP.

We explain the overview and details of our proposed algorithms in Sections 4.1 and 4.2, respectively. Then we analyze the theoretical properties of our algorithms in Section 4.3.

4.1 Overview

Motivation. The drawback of the triangle algorithm in \([30]\) is a prohibitively high download cost at the second round. This comes from the fact that in their algorithm, each user \( v_i \) applies Warner’s RR (Randomized Response) \([61]\) to bits for smaller user IDs in her neighbor list \( a_i \) (i.e., lower triangular part of \( A \)) and then downloads the whole noisy graph. Since Warner’s RR outputs 1 (edge) with high probability (e.g., about 0.5 when \( \varepsilon \) is close to 0), the number of edges in the noisy graph is extremely large—about half of the \( \binom{n}{2} \) possible edges will be edges.

In this paper, we address this issue by introducing two strategies: sampling edges and selecting edges each user downloads. First, each user \( v_i \) samples each 1 (edge) after applying Warner’s RR. Edge sampling has been widely studied in a non-private triangle counting problem \([14, 26, 58, 62]\). In particular, Wu et al. \([62]\) compare various non-private triangle algorithms (e.g., edge sampling, node sampling, triangle sampling) and show that edge sampling provides almost the lowest estimation error. They also formally prove that edge sampling outperforms node sampling. Thus, sampling edges after Warner’s RR is a natural choice for our private setting.

Second, we propose three strategies for selecting edges each user downloads. The first strategy is to simply select all noisy edges; i.e., each user downloads the whole noisy graph in the same way as \([30]\). The second and third strategies select some edges (rather than all edges) in a more clever manner so that the estimation error is significantly reduced. We provide a more detailed explanation in Section 4.2.

Algorithm Overview. Figure 2 shows the overview of our proposed algorithms.

At the first round, each user \( v_i \) obfuscates bits for smaller user IDs in her neighbor list \( a_i \) by an LDP mechanism which we call the ARR (Asymmetric Randomized Response) and
sends the obfuscated bits to a server. The ARR is a combination of Warner’s RR and edge sampling; i.e., we apply Warner’s RR that outputs 1 or 0 as it is with probability \( \rho = \frac{\epsilon}{\pi^2} \) and then sample each 1 with probability \( p_2 \in [0, 1] \). Unlike Warner’s RR, the ARR is asymmetric in that the flip probability in the whole process is different depending on the input value. As with Warner’s RR, the ARR provides edge LDP. We can also significantly reduce the number of 1s (hence the communication cost) by setting \( p_2 \) small.

At the second round, the server calculates a message \( M_i \) for user \( v_i \) consisting of some or all noisy edges between others. We propose three strategies for calculating \( M_i \). User \( v_i \) downloads \( M_i \) from the server. Then, since user \( v_i \) knows her edges, \( v_i \) can count noisy triangles \((v_i, v_j, v_k)\) such that \( j < k < i \) and only one edge \((v_j, v_k)\) is noisy, as shown in Figure 2. The condition \( j < k < i \) is imposed to use only the lower triangular part of \( A \), i.e., to avoid the doubling issue in Section 3.2. User \( v_i \) adds a corrective term and the Laplacian noise to the noisy triangle count and sends it to a server. The corrective term is added to enable the server to obtain an unbiased estimate of \( f_{\Delta}(G) \). The Laplacian noise provides an unbiased estimate with a much smaller variance. See Appendix B for a detailed comparison.

4.2 Algorithms

ARR. First, we formally define the ARR. The ARR has two parameters; \( \varepsilon \in \mathbb{R}_{>0} \) and \( \mu \in [0, \frac{\pi^2}{\epsilon^2}] \). The parameter \( \varepsilon \) is the privacy budget, and \( \mu \) controls the communication cost.

Let \( \text{ARR}_{\varepsilon, \mu} \) be the ARR with parameters \( \varepsilon \) and \( \mu \). It takes 0/1 as input and outputs 0/1 with the following probability:

\[
\Pr[\text{ARR}_{\varepsilon, \mu}(1) = b] = \begin{cases} 
\mu & (b = 1) \\
1 - \mu & (b = 0)
\end{cases}
\]

(5)

\[
\Pr[\text{ARR}_{\varepsilon, \mu}(0) = b] = \begin{cases} 
\mu \rho & (b = 1) \\
1 - \mu \rho & (b = 0)
\end{cases}
\]

(6)

where \( \rho = e^{-\varepsilon} \). By Figure 2, we can view this randomization as a combination of Warner’s RR [61] and edge sampling, where \( \mu = p_1 p_2 \). In fact, the ARR with \( \mu = p_1 = \frac{\epsilon}{\pi^2} \) (i.e., \( p_2 = 1 \)) is equivalent to Warner’s RR.

Each user \( v_i \) applies the ARR to bits for smaller user IDs in her neighbor list \( \mathbf{a}_i \); i.e., \( \mathcal{R}_j(\mathbf{a}_i) = (\text{ARR}_{\varepsilon, \mu}(a_{i,1}), \ldots, \text{ARR}_{\varepsilon, \mu}(a_{i,n-1})) \). Then \( v_i \) sends \( \mathcal{R}_j(\mathbf{a}_i) \) to the server. Since applying Warner’s RR to \( \mathbf{a}_i \) provides \( \varepsilon \)-edge LDP (as described in Section 3.2) and the sampling is a post-processing process, applying the ARR to \( \mathbf{a}_i \) also provides \( \varepsilon \)-edge LDP by the immunity to post-processing [25].

Let \( E' \subseteq V \times V \) be a set of noisy edges sent by users.

Which Noisy Edges to Download? Now, the main question tackled in this paper is: Which noisy edges should each user \( v_i \) download at the second round? Note that user \( v_i \) is not allowed to download only a set of noisy edges that form noisy triangles (i.e., \{ \((v_j, v_k) \in E', (v_i, v_j) \in E, (v_i, v_k) \in E)\} because it tells the server who are friends with \( v_i \). In other words, user \( v_i \) cannot leak her original edges to the server when she downloads noisy edges; the server must choose which part of \( E' \) to include in the message \( M_i \) it sends her.

Thus, a natural solution would be to download all noisy edges between others (with smaller user IDs); i.e., \( M_i = \{(v_j, v_k) \in E' \mid j < k < i\} \). We denote our algorithm with this full download strategy by \( \text{ARRFull}_\Delta \). The (inefficient) two-rounds algorithm in [30] is a special case of \( \text{ARRFull}_\Delta \) without sampling (\( \mu = p_1 \)). In other words, \( \text{ARRFull}_\Delta \) is a generalization of the two-rounds algorithm in [30] using the ARR.

In this paper, we show that we can do much better than \( \text{ARRFull}_\Delta \). Specifically, we prove in Section 4.3 that \( \text{ARRFull}_\Delta \) results in a high estimation error when the number of 4-cycles (cycles of length 4) in \( G \) is large. Intuitively, this can be explained as follows. Suppose that \( v_i, v_j, v_k, \) and \( v_\ell \) (\( j < k < i, j < k < \ell \)) form a 4-cycle. There is no triangle in this graph. However, if there is a noisy edge between \( v_i \) and \( v_k \), then two (incorrect) noisy triangles appear: \((v_i, v_j, v_k)\) counted by \( v_i \) and \((v_\ell, v_j, v_k)\) counted by \( v_\ell \). More generally, let \( E_{jk} \) (resp. \( E_{j\ell k} \)) \( \in \{0, 1\} \) be a random variable that takes 1 if \((v_i, v_j, v_k)\) (resp. \((v_\ell, v_j, v_k)\)) forms a noisy triangle and 0 otherwise. Then, the covariance \( \text{Cov}(E_{jk}, E_{j\ell k}) \) between \( E_{jk} \) and \( E_{j\ell k} \) is large because the presence/absence of a single noisy edge \((v_j, v_k)\) affects the two noisy triangles.

To address this issue, we introduce a trick that makes the two noisy triangles less correlated with each other. We call this the 4-cycle trick. Specifically, we propose two algorithms in which the server uses noisy edges connected to \( v_i \) when it calculates a message \( M_i \) for \( v_i \). In the first algorithm, the server selects noisy edges \((v_j, v_k)\) such that one noisy edge is connected from \( v_k \) to \( v_i \); i.e., \( M_i = \{(v_j, v_k) \in E' \mid (v_j, v_k) \in E', j < k < i\} \). We call this algorithm \( \text{ARROneNS}_\Delta \), as one noisy edge is connected to \( v_i \). In the second algorithm, the server selects noisy edges \((v_j, v_k)\) such that two noisy edges are connected from these nodes to \( v_i \); i.e., \( M_i = \{(v_j, v_k) \in E' \mid (v_j, v_k) \in E', (v_j, v_k) \in E', j < k < i\} \). We call this algorithm \( \text{ARRTwoNS}_\Delta \), as two noisy edges are connected to \( v_i \). Note that user \( v_i \) does not leak her original edges to the server at the time of download in these algorithms, because the server uses only noisy edges \( E' \) sent by users to calculate \( M_i \).

Figure 3 shows our three algorithms. The download cost \( \text{Cost}_{DL} \) in (3) is \( O(\mu^2 \log n), O(\mu^2 n^2 \log n) \), and
Noisy edges between others (v_j and v_k) to DL

Cost_{DL} | O(µ^3 n^2 log n) | O(µ^2 n^2 log n) | O(µ n^2 log n)

Figure 3: Noisy edges to download in our three algorithms.

Figure 4: 4-cycle trick. ARRFull_Δ counts two (incorrect) noisy triangles when one noisy edge appears. ARROneNS_Δ and ARRTwoNS_Δ avoid this by increasing independent noise.

O(µ^3 n^2 log n), respectively, when we regard ε as a constant. In our experiments, we set the parameter µ in the ARR so that µ in ARRFull_Δ is equal to µ^2 in ARROneNS_Δ and also equal to µ^3 in ARRTwoNS_Δ; e.g., µ = 10^{-6}, 10^{-3}, and 10^{-2} in ARRFull_Δ, ARROneNS_Δ, and ARRTwoNS_Δ, respectively. Then the download cost is the same between the three algorithms.

Figure 4 shows our 4-cycle trick. ARRFull_Δ counts two (incorrect) noisy triangles when a noisy edge (v_j, v_k) appears. In contrast, ARROneNS_Δ (resp. ARRTwoNS_Δ) counts both the two noisy triangles only when three (resp. five) independent noisy edges appear, as shown in Figure 4. Thus, this bad event happens with a much smaller probability. For example, ARRFull_Δ (µ = 10^{-6}), ARROneNS_Δ (µ = 10^{-3}), and ARRTwoNS_Δ (µ = 10^{-2}) count both the two noisy triangles with probability 10^{-6}, 10^{-9}, and 10^{-10}, respectively. The covariance Cov(E_{i,j}, E_{j,k}) of ARROneNS_Δ and ARRTwoNS_Δ is also much smaller than that of ARRFull_Δ.

In our experiments, we show that ARROneNS_Δ and ARRTwoNS_Δ significantly outperforms ARRFull_Δ for a large-scale graph or dense graph, in both of which the number of 4-cycles in G is large.

ARROneNS_Δ vs. ARRTwoNS_Δ. One might expect that ARRTwoNS_Δ outperforms ARROneNS_Δ because ARRTwoNS_Δ addresses the 4-cycle issue more aggressively; i.e., the number of independent noisy edges in a 4-cycle is larger in ARRTwoNS_Δ, as shown in Figure 4. However, ARROneNS_Δ can reduce the global sensitivity of the Laplacian noise at the second round more effectively than ARRTwoNS_Δ, as explained in Section 5. Consequently, ARROneNS_Δ, which is the most tricky algorithm, achieves the smallest estimation error in our experiments. See Sections 5 and 6 for details of the global sensitivity and experiments, respectively.

Three Algorithms. Below we explain the details of our three algorithms. For ease of explanation, we assume that the maximum degree d_{max} is public in Section 4.2. Note, however, that our double clipping (which is proposed to significantly reduce the global sensitivity) in Section 5 does not assume that d_{max} is public. Consequently, our entire algorithms do not require the assumption that d_{max} is public.

Recall that the server calculates a message M_i for v_i as:

\[ M_i = \{ (v_j, v_k) \in E \mid j < k < i \} \]  (7)
\[ M_i = (v_j, v_k) \in E \} \]  (8)
\[ M_i = (v_j, v_k) \in E \} \]  (9)
in ARRFull_Δ, ARROneNS_Δ, ARRTwoNS_Δ, respectively.

Algorithm 1 shows our three algorithms. These algorithms are processed differently in lines 2 and 9; “F”, “O”, “T” are shorthands for ARRFull_Δ, ARROneNS_Δ, and ARRTwoNS_Δ,

\[ \hat{f}_{\Delta}(G) = \frac{1}{\rho (1-p)} \sum_{\rho=1}^{\rho} \hat{w}_i; \]

\[ \textbf{Input:} \text{Graph } G \in \mathcal{G} \text{ represented as neighbor lists } \mathbf{a}_1, \ldots, \mathbf{a}_n \in \{0, 1\}^n, \text{ privacy budgets } \mathbf{e}_1, \mathbf{e}_2 \in \mathbb{R}_{\geq 0}, d_{max} \in \mathbb{Z}_{\geq 0}, \mu \in [0, \frac{e_1}{e_1+1}]. \]

\[ \textbf{Output:} \text{Private estimate } f_{\Delta}(G) \text{ of } f_{\Delta}(G). \]

\[ 1 \] [s] \( \mu \leftarrow e^{-\epsilon_1}; \)
\[ 2 \] [v_i, s] \( \mu \leftarrow \mu, \rho; \) and \( \mu \) in F, O, and T, respectively; \( \rho \) / First round. * /
\[ 3 \] for \( i = 1 \) to \( n \) do
\[ 4 \] \( [v_i] r_i \leftarrow (ARR_{\epsilon_{1}}, (a_i, 1, \ldots, ARR_{\epsilon_{1}}, (a_{i-j-1}) \); \)
\[ 5 \] \( [v_i] \text{ Upload } r_i = (r_{i,1}, \ldots, r_{i,d-1}) \) to the server; \( \text{/* Second round. */} \)
\[ 6 \] end
\[ 7 \] [s] \( E' = \{ (v_j, v_k) : r_{k,j} = 1, j < k \}; \)
\[ 8 \] for \( i = 1 \) to \( n \) do
\[ 9 \] [s] Compute \( M_i \) by (7), (8), and (9) in F, O, and T, respectively;
\[ 10 \] [v_i] Download \( M_i \) from the server;
\[ 11 \] [v_i] \( t_i \leftarrow \{(v_j, v_k) : a_{i,j} = a_{i,k} = 1, (v_j, v_k) \in M_i, j < k < i \}; \)
\[ 12 \] [v_i] \( s_i \leftarrow \{(v_j, v_k) : a_{i,j} = a_{i,k} = 1, j < k < i \}; \)
\[ 13 \] [v_i] \( w_i \leftarrow t_i - \rho \Delta_s; \)
\[ 14 \] [v_i] \( \hat{w}_i \leftarrow w_i + \text{Lap} \left( \frac{d_{max}}{\rho} \right); \)
\[ 15 \] [v_i] \text{ Upload } \hat{w}_i \) to the server;
\[ 16 \] end
\[ 17 \] [s] \( \hat{f}_{\Delta}(G) \leftarrow \frac{1}{\rho (1-p)} \sum_{\rho=1}^{\rho} \hat{w}_i; \)
\[ 18 \] return \( \hat{f}_{\Delta}(G) \)

\[ \text{Algorithm 1: Our three algorithms. “F”, “O”, “T” are shorthands for ARRFull_Δ, ARROneNS_Δ, and ARRTwoNS_Δ, respectively. [v_i] and [s] represent that the process is run by v_i and the server, respectively.} \]
respectively. The privacy budgets for the first and second rounds are $\epsilon_1, \epsilon_2 \in \mathbb{R}_{\geq 0}$, respectively.

The first round appears in lines 3-7 of Algorithm 1. In this round, each user applies $ARR_{\epsilon_1, \mu}$ defined by (5) and (6) to bits $a_{i,1}, \ldots, a_{i,j-1}$ for smaller user IDs in her neighbor list $a_i$, i.e., lower triangular part of $A$. Let $v_i = (r_{i,1}, \ldots, r_{i,i-1}) \in \{0, 1\}^{i-1}$ be the obfuscated bits of $v_i$. User $v_i$ uploads $r_i$ to the server. Then the server combines the noisy edges together, forming $E' = \{(v_j, v_k) : r_{i,j} = 1, j < k\}$.

The second round appears in lines 8-17 of Algorithm 1. In this round, the server computes a message $M_i$ by (7), (8), or (9), and user $v_i$ downloads it. Then user $v_i$ calculates the number $t_i \in \mathbb{Z}_{\geq 0}$ of noisy triangles $(v_i, v_j, v_k)$ such that only one edge $(v_j, v_k)$ is noisy, as shown in Figure 2. User $v_i$ also calculates a corrective term $s_i \in \mathbb{Z}_{\geq 0}$. The corrective term $s_i$ is the number of possible triangles involving $v_i$ and is computed to obtain an unbiased estimate of $f_\Delta(G)$. User $v_i$ calculates $w_i = t_i - \mu^2 s_i$, where $\mu = e^{-\epsilon_1}$ and $\mu = \mu^2, e^2$ and $\mu^2$ in “F”, “O”, and “T”, respectively. Then $v_i$ adds the Laplacian noise $\text{Lap}(\frac{d_{\text{max}}}{\epsilon_2})$ to $w_i$ to provide $\epsilon_2$-edge LDP and sends the noisy value $\tilde{w}_i = (w_i + \text{Lap}(\frac{d_{\text{max}}}{\epsilon_2}))$ to the server. Note that adding one edge increases both $t_i$ and $s_i$ by at most $d_{\text{max}}$. Thus, the global sensitivity of $w_i$ is at most $d_{\text{max}}$. Finally, the server calculates an estimate of $f_\Delta(G)$ as: $f_\Delta(G) = \frac{1}{\mu(1-\mu)} \sum_{i=1}^n \tilde{w}_i$.

As we prove later, $f_\Delta(G)$ is an unbiased estimate of $f_\Delta(G)$.

4.3 Theoretical Analysis

We now introduce the theoretical guarantees on the privacy, communication, and utility of our algorithms.

Privacy. We first show the privacy guarantees:

**Theorem 1.** For $i \in [n]$, let $R_{\epsilon_1}^1, R_{\epsilon_2}^2(M_i)$ be the randomizers used by user $v_i$ in rounds 1 and 2 of Algorithm 1. Let $R_{\epsilon_1}^0(a) = (R_{\epsilon_1}^1(a), R_{\epsilon_2}^2(M_i)(a))$ be the composition of the two randomizers. Then, $R_{\epsilon_1}$ satisfies $(\epsilon_1 + \epsilon_2)$-edge LDP and $(R_{\epsilon_1}, \ldots, R_{\epsilon_n})$ satisfies $(\epsilon_1 + \epsilon_2)$-relationship DP.

Note that the doubling issue in Section 3.2 does not occur, because we use only the lower triangular part of $A$. By the immunity to post-processing, the estimate $\hat{f}_\Delta(G)$ also satisfies $(\epsilon_1 + \epsilon_2)$-edge LDP and $(\epsilon_1 + \epsilon_2)$-relationship DP.

Communication. Recall that we evaluate the algorithms based on their download cost (3) and upload cost (4).

**Download Cost:** The download cost is the number of bits required to download $M_i$. $M_i$ can be represented as a list of edges between others, and each edge can be identified with two indices (user IDs), i.e., $2 \log n$ bits. There are $\frac{\binom{n-1}{2}}{2} \approx \frac{n^2}{2}$ edges between others. $ARR_{\epsilon_1, \mu}$ outputs 1 with probability at most $\mu$. In addition, each noisy triangle must have 1, 2, and 3 noisy edges in $ARR_{\text{Full}}, ARR_{\text{OneNS}},$ and $ARR_{\text{TwoNS}}$, respectively, as shown in Figure 3.

Thus, the download cost in Algorithm 1 can be written as:

$$\text{Cost}_{DL} \leq \mu n^2 \log n,$$

where $\mu = \mu^2$, and $\mu^3$ in $ARR_{\text{Full}}, ARR_{\text{OneNS}},$ and $ARR_{\text{TwoNS}}$, respectively. In (10), we upper-bounded $\text{Cost}_{DL}$ by using the fact that $ARR_{\epsilon_1, \mu}$ outputs 1 with probability at most $\mu$. However, when $d_{\text{max}} \ll n$, $ARR_{\epsilon_1, \mu}$ outputs 1 with probability $\mu e^{-\epsilon_1}$ in most cases. In that case, we can roughly approximate $\text{Cost}_{DL}$ by replacing $\mu$ with $\mu e^{-\epsilon_1}$ in (10).

**Upload Cost:** The upload cost comes from the number of bits required to upload $R_{\epsilon_1}^1(a)$ and $R_{\epsilon_2}^2(M_i)(a)$. Uploading $R_{\epsilon_1}^1(a)$ involves uploading $r_i$ (line 5), which is a list of up to $n$ noisy neighbors. By sending just the indices (user IDs) of the $1$s in $r_i$, each user sends $\|r_i\|_1 \log n$ bits, where $\|r_i\|_1$ is the number of $1$s in $r_i$. When we use $ARR_{\epsilon_1, \mu}$, we have $\|r_i\|_1 \leq \mu n$. Uploading $R_{\epsilon_2}^2(M_i)$ involves uploading a single real number $\tilde{w}_i$ (line 15), which is negligibly small (e.g., 64 bits when we use a double-precision floating-point). Thus, the upload cost in Algorithm 1 can be written as:

$$\text{Cost}_{UL} \leq \mu n \log n.$$
5 Double Clipping

In Section 4, we showed that the estimation error caused by empirical estimation (i.e., the first term in Theorem 2) is significantly reduced by the 4-cycle trick. However, the estimation error is still very large in our algorithms presented in Section 4, as shown in our experiments. This is because the estimation error by the Laplacian noise (i.e., the second term in Theorem 2) is very large, especially for small $\varepsilon_2$ or $\mu$. This error term is tight and unavoidable as long as we use $d_{max}$ as a global sensitivity, which suggests that we need a better global sensitivity analysis. To significantly reduce the global sensitivity, we propose a novel double clipping technique.

We describe the overview and details of our double clipping in Sections 5.1 and 5.2, respectively. Then we perform theoretical analysis in Section 5.3.

5.1 Overview

Motivation. Figure 5 shows noisy triangles involving edge $(v_i, v_j)$ counted by user $v_i$ in our three algorithms. Our algorithms in Section 4 use the fact that the number of such noisy triangles (hence the global sensitivity) is upper-bounded by the maximum degree $d_{max}$ because adding one edge increases the triangle count by at most $d_{max}$. Unfortunately, this upper-bound is too large, as shown in our experiments.

In this paper, we significantly reduce this upper-bound by using the parameter $\mu$ in the ARR and user $v_i$’s degree $d_i \in \mathbb{Z}_{\geq 0}$ for users with smaller IDs. For example, the number of noisy triangles involving $(v_i, v_j)$ in ARRFull is expected to be around $\mu d_i$ because one noisy edge is included in each noisy triangle (as shown in Figure 5) and all noisy edges are independent. $\mu d_i$ is very small, especially when we set $\mu \ll 1$ to reduce the communication cost.

However, we cannot directly use $\mu d_i$ as an upper-bound of the global sensitivity in ARRFull for two reasons. First, $\mu d_i$ leaks the exact value of user $v_i$’s degree $d_i$ and violates edge LPD. Second, the number of noisy triangles involving $(v_i, v_j)$ exceeds $\mu d_i$ with high probability (about 0.5). Thus, the noisy triangle count cannot be upper-bound by $\mu d_i$.

To address these two issues, we propose a double clipping technique, which is explained below.

Algorithm Overview. Figure 6 shows the overview of our double clipping, which consists of an edge clipping and noisy triangle clipping. The edge clipping addresses the first issue (i.e., leakage of $d_i$) as follows. It privately computes a noisy version of $d_i$ (denoted by $\hat{d}_i$) with edge LPD. Then it removes some neighbors from a neighbor list $a_i$ so that the degree of $v_i$ never exceeds the noisy degree $\hat{d}_i$. This removal process is also known as graph projection [23, 24, 37, 50]. Edge clipping is used in [30] to obtain a noisy version of $d_{max}$.

The main novelty in our double clipping lies at the noisy triangle clipping to address the second issue (i.e., excess of the noisy triangle count). This issue appears when we attempt to reduce the global sensitivity by using a very small sampling probability for each edge. Therefore, the noisy triangle clipping has not been studied in the existing works on private triangle counting [24, 30, 36, 37, 56, 64, 65, 67], because they do not apply a sampling technique.

Our noisy triangle clipping reduces the noisy triangle count so that it never exceeds a user-dependent clipping threshold $\kappa_i \in \mathbb{R}_{\geq 0}$. Then a crucial issue is how to set an appropriate threshold $\kappa_i$. We theoretically analyze the probability that the noisy triangle count exceeds $\kappa_i$ (referred to as the triangle excess probability) as a function of the ARR parameter $\mu$ and the noisy degree $\hat{d}_i$. Then we set $\kappa_i$ so that the triangle excess probability is very small ($= 10^{-5}$ in our experiments).

We use the clipping threshold $\kappa_i$ as a global sensitivity. Note that $\kappa_i$ provides edge LPD because $\hat{d}_i$ provides edge LPD, i.e., immunity to post-processing [25]. $\kappa_i$ is also very small when $\mu \ll 1$, as it is determined based on $\mu$.

5.2 Algorithms

Algorithm 2 shows our double clipping algorithm. All the processes are run by user $v_i$ at the second round. Thus, there is no interaction with the server in Algorithm 2.

Edge Clipping. The edge clipping appears in lines 2-3 of Algorithm 2. It uses a privacy budget $\varepsilon_0 \in \mathbb{R}_{\geq 0}$.

In line 2, user $v_i$ adds the Laplacian noise $\text{Lap}(\frac{1}{\varepsilon_0})$ to her degree $d_i$. Since adding/removing one edge changes $d_i$ by at most 1, this process provides $\varepsilon_0$-edge LPD. $v_i$ also adds some non-negative constant $\alpha \in \mathbb{R}_{\geq 0}$ to $d_i$. We add this value so that edge removal (in line 3) occurs with a very small probability; e.g., in our experiments, we set $\alpha = 150$, where edge removal occurs with probability $1.5 \times 10^{-7}$ when $\varepsilon_0 = 0.1$. A similar technique is introduced in [56] to provide $(\varepsilon, \delta)$-DP [25] with small $\delta$. The difference between ours and [56] is that we perform edge clipping to always provide $\varepsilon$-DP; i.e., $\delta = 0$. Let $\hat{d}_i \in \mathbb{R}_{\geq 0}$ be the noisy degree of $v_i$. 
Algorithm 2: Our double clipping algorithm. “F”, “O”, “T” are shorthands for ARRFull$\Delta$, ARROneNS$\Delta$, and ARRTwoNS$\Delta$, respectively. All the processes are run by user $v_i$.

In line 3, user $v_i$ calls the function $\text{GraphProjection}$, which performs graph projection as follows: if $d_i > \hat{d}_i$, randomly remove $d_i - \hat{d}_i$ neighbors from $a_i$; otherwise, do nothing. Consequently, the degree of $v_i$ never exceeds $\hat{d}_i$.

Noisy Triangle Clipping. The noisy triangle clipping appears in lines 4-11 of Algorithm 2.

In lines 4-6, user $v_i$ calculates the number $t_{i,j}$ in $\mathbb{Z}_{\geq 0}$ of noisy triangles $(v_i, v_j, v_k)$ $(j < k < i)$ involving $(v_i, v_j)$ (as shown in Figure 5). Note that the total number $t_i$ of noisy triangles of $v_i$ can be expressed as: $t_i = \sum_{j=1, k<i} t_{i,j}$. In line 7, $v_i$ calls the function $\text{ClippingThreshold}$, which calculates a clipping threshold $\kappa_i \in [\mu d_i, \hat{d}_i]$ ($\mu = \mu, \mu^2$, and $\mu^3$ in “F”, “O”, and “T”, respectively) based on the ARR parameter $\mu$ and the noisy degree $\hat{d}_i$ so that the triangle excess probability does not exceed some constant $\beta \in \mathbb{R}_{\geq 0}$. We explain how to calculate the triangle excess probability in Section 5.3. In line 8, $v_i$ calculates the total number $t_i$ of noisy triangles by summing up $t_{i,j}$ with the exception that $v_i$ adds $\kappa_i$ if $t_{i,j} > \kappa_i$. In other words, triangle removal occurs if $t_{i,j} > \kappa_i$. Then, the number of noisy triangles involving $(v_i, v_j)$ never exceeds $\kappa_i$.

Lines 9-11 in Algorithm 2 are the same as lines 12-14 in Algorithm 1, except that the global sensitivity in the former (resp. latter) is $\kappa_i$ (resp. $d_{\text{max}}$). Line 11 in Algorithm 2 provides $\varepsilon_2$-edge LDP because the number of triangles involving $(v_i, v_j)$ is now upper-bounded by $\kappa_i$.

Our Entire Algorithms with Double Clipping. We can run our algorithms $\text{ARRFull}_\Delta$, $\text{ARROneNS}_\Delta$, $\text{ARRTwoNS}_\Delta$ with double clipping just by replacing lines 11-14 in Algorithm 1 with lines 2-11 in Algorithm 2. That is, after calculating $\hat{w}_i$ by Algorithm 2, $v_i$ uploads $\hat{w}_i$ to the server. Then the server calculates an estimate of $f_{\Delta}(G)$ as $\hat{f}_{\Delta}(G) = \frac{1}{p(1-p)} \sum_{i=1}^{n} \hat{w}_i$.

We also note that the input $d_{\text{max}}$ in Algorithm 1 is no longer necessary thanks to the edge clipping; i.e., our entire algorithms with double clipping do not assume that $d_{\text{max}}$ is public.

5.3 Theoretical Analysis

We now perform a theoretical analysis on the privacy and utility of our double clipping.

Privacy. We begin with the privacy guarantees:

Theorem 3. For $i \in [n]$, let $\mathcal{R}_i^1, \mathcal{R}_i^2 (\mathcal{M}_i)$ be the randomizers used by user $v_i$ in rounds 1 and 2 of our algorithms with double clipping (Algorithms 1 and 2). Let $\mathcal{R}_i (a_i) = (\mathcal{R}_i^1 (a_i), \mathcal{R}_i^2 (\mathcal{M}_i (a_i)))$ be the composition of the two randomizers. Then, $\mathcal{R}_i$ satisfies $(\varepsilon_0 + \varepsilon_1 + \varepsilon_2)^3$-edge LDP, and $(\mathcal{R}_1, \ldots, \mathcal{R}_n)$ satisfies $(\varepsilon_0 + \varepsilon_1 + \varepsilon_2)$-relationship DP.

Utility. Next, we show the triangle excess probability:

Theorem 4. In Algorithm 2, the triangle excess probability (i.e., probability that the number of noisy triangles $t_{i,j}$ involving edge $(v_i, v_j)$ exceeds a clipping threshold $\kappa_i$) is:

$$\Pr(t_{i,j} > \kappa_i) \leq \exp \left[ -\hat{d}_i D \left( \frac{\kappa_i}{\hat{d}_i} \parallel \mu \right) \right]$$

in $\text{ARRFull}_\Delta$, $\text{ARROneNS}_\Delta$, and $\text{ARRTwoNS}_\Delta$, respectively, where $D(p_1 \parallel p_2)$ is the Kullback-Leibler divergence between two Bernoulli distributions; i.e.,

$$D(p_1 \parallel p_2) = p_1 \log \frac{p_1}{p_2} + (1-p_1) \log \frac{1-p_1}{1-p_2}.$$
(14) are $2.5 \times 10^{-12}, 2.5 \times 10^{-12},$ and $3.3 \times 10^{-2},$ respectively. Consequently, ARRTwoNS\(_\triangle\) has a larger global sensitivity \(\kappa_i\) for the same value of \(\beta\).

We can explain a large global sensitivity \(\kappa_i\) of ARRTwoNS\(_\triangle\) as follows. The number \(t_{i,j}\) of noisy triangles involving \((v_i, v_j)\) in ARRFull\(_\triangle\) is expected to be around \(\mu d_i\) because one noisy edge is in each noisy triangle (as in Figure 5) and all noisy edges are independent. For the same reason, \(t_{i,j}\) in ARROneNS\(_\triangle\) is expected to be around \(\mu^2 d_i\). However, \(t_{i,j}\) in ARRTwoNS\(_\triangle\) is not expected to be around \(\mu^2 d_i\), because all the noisy triangles have noisy edge \((v_i, v_j)\) in common (as in Figure 5). Then, the expectation of \(t_{i,j}\) largely depends on the presence/absence of the noisy edge \((v_i, v_j)\); i.e., if noisy edge \((v_i, v_j)\) exists, it is \(\mu^2 d_i\); otherwise, 0. Thus, \(\kappa_i\) cannot be effectively reduced by double clipping.

**Summary.** The performance guarantees of our three algorithms with double clipping can be summarized in Table 1.

The first and second terms of the expected \(l_2\) loss are the \(l_2\) loss of empirical estimation and that of the Laplacian noise, respectively. For small \(\mu\), the \(l_2\) loss of empirical estimation can be expressed as \(O(n d_{\max}^3), O(n d_{\max}^2),\) and \(O(n^2 d_{\max})\) in ARRFull\(_\triangle\), ARROneNS\(_\triangle\), and ARRTwoNS\(_\triangle\), respectively, as explained in Section 4.3. The \(l_2\) loss of the Laplacian noise is \(O(\sum_{i=1}^n \kappa_i^2)\), which is much smaller than \(O(n d_{\max}^3)\). Thus, our ARROneNS\(_\triangle\) that effectively reduces \(\kappa_i\) provides the smallest error, as shown in our experiments.

We also note that both the space and the time complexity to compute and send \(M_i\) in our algorithms are \(O(\mu^2 n^2)\) (as \(|E'| = O(\mu^2 n^2)\)), which is much smaller than [30] (= \(O(n^2)\)).

### 6 Experiments

To evaluate each component of our algorithms in Sections 4 and 5 as well as our entire algorithms (i.e., ARRFull\(_\triangle\), ARROneNS\(_\triangle\), and ARRTwoNS\(_\triangle\) with double clipping), we pose the following three research questions:

**RQ1.** How do our three triangle counting algorithms (i.e., ARRFull\(_\triangle\), ARROneNS\(_\triangle\), and ARRTwoNS\(_\triangle\)) in Section 4 compare with each other in terms of accuracy?

**RQ2.** How much does our double clipping technique in Section 5 decrease the estimation error?

**RQ3.** How much do our entire algorithms reduce the communication cost, compared to the existing algorithm [30], while keeping high utility (e.g., relative error \(\ll 1\))? In Appendix B, we also compare our entire algorithms with one-round algorithms.

### 6.1 Experimental Set-up

In our experiments, we used two real graph datasets:

**Gplus.** The Google+ dataset [41] (denoted by Gplus) was collected from users who had shared circles. From the dataset, we constructed a social graph \(G = (V, E)\) with 107614 nodes (users) and 12238285 edges, where edge \((v_i, v_j) \in E\) represents that \(v_i\) follows or is followed by \(v_j\). The average (resp. maximum) degree in \(G\) is 113.7 (resp. 20127).

**IMDB.** The IMDB (Internet Movie Database) [2] (denoted by IMDB) includes a bipartite graph between 896308 actors and 428440 movies. From this, we constructed a graph \(G = (V, E)\) with 896308 nodes (actors) and 57064358 edges, where edge \((v_i, v_j) \in E\) represents that \(v_i\) and \(v_j\) have played in the same movie. The average (resp. maximum) degree in \(G\) is 63.7 (resp. 15451). Thus, IMDB is more sparse than Gplus.

In Appendix D, we also evaluate our algorithms using a synthetic graph based on the Barabási-Albert model [11], which has a power-law degree distribution.

We evaluated our algorithms while changing \(\mu^*\), where \(\mu^* = \mu, \mu^2,\) and \(\mu^3\) in ARRFull\(_\triangle\), ARROneNS\(_\triangle\), and ARRTwoNS\(_\triangle\), respectively. Cost\(_{DL}\) is the same between the three algorithms. We typically set the total privacy budget \(\epsilon\) to \(\epsilon = 1\) (at most 2) because it is acceptable in many practical scenarios [39].

In our double clipping, we set \(\alpha = 150\) and \(\beta = 10^{-6}\) so that both edge removal and triangle removal occur with a very small probability \((\leq 10^{-6}\) when \(\epsilon_0 = 0.1\)). Then for each algorithm, we evaluated the relative error between the true triangle count \(f_\triangle(G)\) and its estimate \(\hat{f}_\triangle(G)\). Since the estimate \(\hat{f}_\triangle(G)\) varies depending on the randomness of LDP mechanisms, we ran each algorithm \(\tau \in \mathbb{N}\) times (\(\tau = 20\) and 10 for Gplus and IMDB, respectively) and averaged the relative error over the \(\tau\) cases.
6.2 Experimental Results

Performance Comparison. First, we evaluated our algorithms with the Laplacian noise. Specifically, we evaluated all possible combinations of our three algorithms with and without our double clipping (six combinations in total) and compared them with the existing two-rounds algorithm in [30]. For algorithms with double clipping, we divided the total privacy budget $\epsilon$ as: $\epsilon_0 = \frac{\epsilon}{5}$ and $\epsilon_1 = \epsilon_2 = \frac{\epsilon}{3}$. Here, we set a very small budget ($\epsilon_0 = \frac{\epsilon}{10}$) for edge clipping because the degree has a small sensitivity (sensitivity $= 1$). For algorithms without double clipping, we divided $\epsilon$ as $\epsilon_1 = \epsilon_2 = \frac{\epsilon}{4}$ and used the maximum degree $d_{\max}$ as the global sensitivity.

Figures 7 and 8 show the results. Figure 7 highlights the relative error of our three algorithms with double clipping when $\epsilon = 1$ or 2 and $\mu^* = 10^{-3}$ (n = 107614 in Gplus, n = 896308 in IMDB). Figure 8 shows the relative error of our three algorithms with double clipping (“DC”) or without (“d_{max}”) double clipping (n = 107614 in Gplus, n = 896308 in IMDB). RRFull_{\Delta}(d_{\max}) is the algorithm in [30]. Cost_{DL} is an upper-bound in (10). When $\mu^* \geq 0.1$, Cost_{DL} can be 6 Gbits and 400 Gbits in Gplus and IMDB, respectively, by downloading only 0/1 for each pair of users $(v_j, v_k)$.

Communication Cost. From Figure 8 (c) and (d), we can explain how much our algorithms can reduce the download cost while keeping high utility, e.g., relative error $\ll 1$.

For example, when we use the algorithm in [30], the download cost is Cost_{DL} = 400 Gbits in IMDB. Thus, when the download speed is 20 Mbps (recommended speed in YouTube [7]), every user $v_i$ needs 6 hours to download the message $M_i$, which is far from practical. In contrast, our ARROneNS_{\Delta} (DC) can reduce it to 160 Mbits (8 seconds when 20 Mbps download rate) or less while keeping relative error $= 0.21$, which is practical and a dramatic improvement over [30]. We also note that since $d_{\max} \ll n$ in IMDB, Cost_{DL} of our ARROneNS_{\Delta} (DC) can also be roughly approximated by 60 Mbits (3 seconds) by replacing $\mu$ with $\mu e^{-\epsilon_1}$ in (10).

4-Cycle Trick. We also investigated the effectiveness of our 4-cycle trick in ARROneNS_{\Delta} and ARRTwoNS_{\Delta} in detail. To this end, we evaluated our three algorithms when we did not add the Laplacian noise at the second round. Note that they do not provide edge LDP, as $e_2 = \infty$. The purpose here is to purely investigate the effectiveness of the 4-cycle trick related to our first research question RQ1.

Figure 9 shows the results, where $\epsilon_1$ and $\mu^*$ are changed to various values. Figure 9 shows that ARROneNS_{\Delta} and ARRTwoNS_{\Delta} significantly outperform RRFull_{\Delta} when $\mu^*$ is small.
This is because in both ARROneNS and ARRTwoNS, the factors of $C_4$ (#4-cycles) and $S_3$ (#3-stars) in the expected $l_2$ loss diminish for small $\mu$, as explained in Section 4.3. In other words, ARROneNS and ARRTwoNS effectively address the 4-cycle issue. Figure 9 also shows that ARRTwoNS slightly outperforms ARROneNS when $\mu^*$ is small. This is because the factors of $C_4$ and $S_3$ diminish more rapidly; i.e., ARRTwoNS addresses the 4-cycle issue more aggressively.

However, when we add the Laplacian noise, ARRTwoNS (DC) is outperformed by ARROneNS (DC), as shown in Figure 8. This is because ARRTwoNS cannot effectively reduce the global sensitivity by double clipping. In Figure 8, the difference between ARROneNS (DC) and ARRFull (DC) is also small for very small $\varepsilon$ or $\mu^*$ (e.g., $\varepsilon = 0.1, \mu^* = 10^{-6}$) because the Laplacian noise is dominant in this case.

Changing $n$. We finally evaluated our three algorithms with double clipping while changing the number $n$ of users. In both Gplus and IMDB, we randomly selected $n$ users from all users and extracted a graph with $n$ users. Then we evaluated the relative error while changing $n$ to various values.

Figure 10 shows the results, where $\varepsilon = 1 (\varepsilon_0 = 0.1, \varepsilon_1 = \varepsilon_2 = 0.45)$ and $\mu^* = 10^{-3}$. In all three algorithms, the relative error decreases with increase in $n$. This is because the expected $l_2$ loss can be expressed as $O(nd_{max}^3)$ or $O(nd_{max}^2)$ in these algorithms as shown in Section 5.3 and the square of the true triangle count can be expressed as $\Omega(n^2)$. In other words, when $d_{max} \ll n$, the relative error becomes smaller for larger $n$. Figure 10 also shows that for small $n$, ARRTwoNS provides the worst performance and ARROneNS performs almost the same as ARRFull. For large $n$, ARRFull performs the worst and ARROneNS performs the best.

To investigate the reason for this, we decomposed the estimation error into two components — the first error caused by empirical estimation and the second error caused by the Laplacian noise. Specifically, for each algorithm, we evaluated the first error by calculating the relative error when we did not add the Laplacian noise ($\varepsilon_1 = 0.45$). Then we evaluated the second error by subtracting the first error from the relative error when we used double clipping ($\varepsilon_0 = 0.1, \varepsilon_1 = \varepsilon_2 = 0.45$).

Figure 11 shows the results for some values of $n$, where “emp” represents the first error by empirical estimation and “Lap” represents the second error by the Laplacian noise. We observe that the second error rapidly decreases with increase in $n$. In addition, the first error of ARRFull is much larger than those of ARROneNS and ARRTwoNS when $n$ is large.

We also examined the number $C_4$ of 4-cycles as a function of $n$. Figure 12 shows the results. We observe that $C_4$ (which is $O(nd_{max}^3)$) is quartic in $n$; e.g., $C_4$ is increased by $2^4 \approx 10$ and $6^4 \approx 10^4$ when $n$ is multiplied by 2 and 6, respectively. This is because we randomly selected $n$ users from all users and $d_{max}$ is almost proportional to $n$ (though $d_{max} \ll n$).

Based on Figures 11 and 12, we can explain Figure 10 as follows. As shown in Section 5.3, the $l_2$ loss of empirical estimation can be expressed as $O(nd_{max}^3), O(nd_{max}^2)$, and
$O(nd_{\text{max}}^2)$ in ARRFull$_\Delta$, ARROneNS$_\Delta$, and ARRTwoNS$_\Delta$, respectively. The large $l_2$ loss of ARRFull$_\Delta$ is caused by a large value of $C_4$. The expected $l_2$ loss of the Laplacian noise is $O(\sum_{i=1}^{n} \kappa_i^2)$, which is much smaller than $O(nd_{\text{max}}^2)$. Thus, as $n$ increases, the Laplacian noise becomes relatively very small, as shown in Figure 11. Consequently, ARROneNS$_\Delta$ provides the best performance for large $n$ because it addresses the 4-cycle issue and effectively reduces the global sensitivity. This explains the results in Figure 10. It is also interesting that when $n \approx 10^5$, ARRFull$_\Delta$ performs the worst in Gplus and almost the same as ARROneNS$_\Delta$ in IMDB (see Figure 10). This is because Gplus is more dense than IMDB and $C_4$ is much larger in Gplus when $n \approx 10^5$, as in Figure 12.

In other words, Figures 10, 11, and 12 are consistent with our theoretical results in Section 5.3. From these results, we conclude that ARROneNS$_\Delta$ is effective especially for a large graph (e.g., $n \approx 10^5$) or dense graph (e.g., Gplus) where the number $C_4$ of 4-cycles is large.

**Summary.** In summary, our answers to our three research questions RQ1-3 are as follows. [RQ1]: Our ARROneNS$_\Delta$ achieves almost the smallest estimation error in all cases and outperforms the other two, especially for a large graph or dense graph where $C_4$ is large. [RQ2]: Our double clipping reduces the estimation error by two or three orders of magnitude. [RQ3]: Our entire algorithm (ARROneNS$_\Delta$ with double clipping) dramatically reduces the communication cost, e.g., from 6 hours to 8 seconds or less (relative error = 0.21) in IMDB at a 20 Mbps download rate [7].

Thus, triangle counting under edge LDP is now much more practical. In Appendix C, we show that the clustering coefficient can also be accurately estimated using our algorithms.

### 7 Conclusions

We proposed triangle counting algorithms under edge LDP with a small estimation error and small communication cost. We showed that our entire algorithms with the 4-cycle trick and double clipping can dramatically reduce the download cost of [30] (e.g., from 6 hours to 8 seconds or less).

We assumed that each user $v_i$ honestly inputs her neighbor list $a_i$, as in most previous work on LDP. However, recent studies [17, 20] show that the estimate in LDP can be skewed by data poisoning attacks. As future work, we would like to analyze the impact of data poisoning on our algorithms and develop defenses (e.g., detection) against it.

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A Basic Notations

Table 2 shows the basic notations used in this paper.

B Comparison with One-Round Algorithms

Below we show that one-round triangle counting algorithms suffer from a prohibitively large estimation error.

First, we note that all of the existing one-round triangle algorithms in [30,64,65] are inefficient and cannot be directly applied to a large-scale graph such as Gplus and IMDB in Section 6. Specifically, in their algorithms, each user $v_i$ applies Warner’s RR to each bit of her neighbor list $a_i$ and sends the noisy neighbor list to the server. Then the server counts...
the number of noisy triangles, each of which has three noisy edges, and estimates $f_{\Delta}(G)$ based on the noisy triangle count.

The noisy graph $G'$ in the server is dense, and there are $O(n^3)$ noisy triangles in $G$. Thus, the time complexity of the existing one-round algorithms [30, 64, 65] is $O(n^7)$. It is also reported in [30] that when $n = 10^6$, the one-round algorithms would require about 35 years even using a supercomputer, due to the enormous number of noisy triangles.

Therefore, we evaluated the existing one-round algorithms by taking the following two steps. First, we evaluate all the existing algorithms in [30, 64, 65] using small graph datasets ($n = 10000$) and show that the algorithm in [30] achieves the lowest estimation error. Second, we improve the time complexity of the algorithm in [30] using the ARR (i.e., edge sampling after Warner’s RR) and compare it with our two-rounds algorithms using large graph datasets in Section 6.

**Small Datasets.** We first evaluated the existing algorithms in [30, 64, 65] using small datasets. For both Gplus and IMDB in Section 6, we first randomly selected $n = 10000$ users from all users and extracted a graph with $n$ users. Then we evaluated the relative error of the following three algorithms: (i) RR (biased) [30, 64], (ii) RR (bias-reduced) [65], and (iii) RR (unbiased) [30]. All of them provide $\varepsilon$-edge LDP.

RR (biased) simply uses the number of noisy triangles in the noisy graph $G'$ obtained by Warner’s RR as an estimate of $f_{\Delta}(G)$. Clearly, it suffers from a very large bias, as $G'$ is dense. RR (bias-reduced) reduces this bias by using a noisy degree sent by each user. However, it introduces some approximation to estimate $f_{\Delta}(G)$, and consequently, it is not clear whether the estimate is unbiased. We used the mean of the noisy degrees as a representative degree to obtain the optimal privacy budget allocation (see [65] for details). RR (unbiased) calculates an unbiased estimate of $f_{\Delta}(G)$ via empirical estimation. It is proved that the estimate is unbiased [30].

In all of the three algorithms, each user $v_i$ obfuscates bits for smaller user IDs in her neighbor list $a_i$. We averaged the relative error over 10 runs.

Figure 13 shows the results. RR (bias-reduced) significantly outperforms RR (biased) and is significantly outperformed by RR (unbiased). We believe this is caused by the fact that RR (bias-reduced) introduces some approximation and does not calculate an unbiased estimate of $f_{\Delta}(G)$.

![Figure 13: Relative error of one-round algorithms for small datasets ($n = 10000$).](image)

**Large Datasets.** Based on Figure 13, we improve the time complexity of RR (unbiased) using the ARR and compare it with our two-rounds algorithms in large datasets.

Specifically, RR (unbiased) counts triangles, 2-edges (three nodes with two edges), 1-edges (three nodes with one edge), and no-edges (three nodes with no edges) in $G'$ obtained by Warner’s RR. Let $m_3, m_2, m_1, m_0 \in \mathbb{Z}_{\geq 0}$ be the numbers of triangles, 2-edges, 1-edges, and no-edges, respectively, after applying Warner’s RR. RR (unbiased) calculates an unbiased estimate of $f_{\Delta}(G)$ from these four values. Thus, we improve RR (unbiased) by using the ARR, which samples each edge with probability $p_2$ after Warner’s RR, and then calculating unbiased estimates of $m_3, m_2, m_1$, and $m_0$.

Let $\hat{m}_3, \hat{m}_2, \hat{m}_1, \hat{m}_0 \in \mathbb{R}$ be the unbiased estimates of $m_3, m_2, m_1$, and $m_0$, respectively. Let $m_3^*, m_2^*, m_1^*, m_0^* \in \mathbb{Z}_{\geq 0}$ be the number of triangles, 2-edges, 1-edges, no-edges, respectively, after applying the ARR. Since the ARR samples each edge with probability $p_2$, we obtain:

\[
\begin{align*}
    m_3^* &= p_2^3 \hat{m}_3 \\
    m_2^* &= 3p_2^2(1-p_2)\hat{m}_3 + p_2^2\hat{m}_2 \\
    m_1^* &= 3p_2(1-p_2)^2\hat{m}_3 + 2p_2(1-p_2)\hat{m}_2 + p_2\hat{m}_1.
\end{align*}
\]
By these equations, we obtain:

\[
\hat{m}_3 = \frac{m_1^3}{p_2^2} \tag{15}
\]

\[
\hat{m}_2 = \frac{m_1^2}{p_2^2} - 3(1 - p_2)\hat{m}_3 \tag{16}
\]

\[
\hat{m}_1 = \frac{m_1}{p_2^2} - 3(1 - p_2)^2\hat{m}_3 - 2(1 - p_2)\hat{m}_2 \tag{17}
\]

\[
\hat{m}_0 = \frac{n(n-1)(n-2)}{6} - \hat{m}_3 - \hat{m}_2 - \hat{m}_1. \tag{18}
\]

Therefore, after applying the ARR to the lower triangular part of A, the server counts \(m_3, m_2, m_1,\) and \(m_0\) in \(G',\) and then calculates the unbiased estimates \(\hat{m}_3, \hat{m}_2, \hat{m}_1,\) and \(\hat{m}_0\) by (15), (16), (17), and (18), respectively. Finally, the server estimates \(f_\Delta(G)\) from \(\hat{m}_3, \hat{m}_2, \hat{m}_1,\) and \(\hat{m}_0\) in the same way as RR (unbiased). We denote this algorithm by ARR (unbiased).

The time complexity of ARR (unbiased) is \(O(\mu^3 n^3)\), where \(\mu\) is the ARR parameter.

We compared ARR (unbiased) with our three algorithms with double clipping using Gplus \((n = 107614)\) and IMDB \((n = 896308)\). For the sampling probability \(p_2,\) we set \(p_2 = 10^{-3}\) or \(10^{-6}\). We averaged the relative error over 10 runs.

Figure 14 shows the results, where we set \(\mu^* = 10^{-6}\) or \(10^{-3}\). In ARR (unbiased), we used \(\mu^*\) as the ARR parameter \(\mu.\) Thus, we can see how much the relative error is reduced by introducing an additional round with ARRFull\(\Delta.\) Figure 14 shows that the relative error of ARR (unbiased) is prohibitively large; i.e., relative error \(\gg 1\). This is because three edges are noisy in any noisy triangle. The relative error is significantly reduced by introducing an additional round because only one edge is noisy in each noisy triangle at the second round.

In summary, one-round algorithms are far from acceptable in terms of the estimation error for large graphs, and two-round algorithms such as ours are necessary.

### C Clustering Coefficient

Here we evaluate the estimation error of the clustering coefficient using our algorithms.

We first estimated a triangle count by using our ARRone\(\Delta\) with double clipping \((\epsilon_0 = \frac{\epsilon}{m}, \epsilon_1 = \frac{\epsilon}{10m})\) because it provides the best performance in Figures 7, 8, and 10. Then we estimated a 2-star count by using the one-round 2-star algorithm in [30] with the edge clipping in Section 5.

Specifically, we calculated a noisy degree \(\hat{d}_i\) of each user \(v_i\) by using the edge clipping with the privacy budget \(\epsilon_0.\) Then we calculated the number \(r_i \in \mathbb{Z}_{\geq 0}\) of 2-stars of which user \(v_i\) is a center, and added \(\text{Lap}(\frac{\Delta}{r_i})\) to \(r_i,\) where \(\Delta = \hat{d}_i.\)

Let \(\hat{r}_i = r_i + \text{Lap}(\hat{d}_i)\) be the noisy 2-star of \(v_i.\) Finally, we calculated the sum \(\sum_{i=1}^n \hat{r}_i\) as an estimate of the 2-star count. This 2-star algorithm provides \((\epsilon_0 + \epsilon_1)\)-edge privacy (see [30] for details). For the privacy budgets \(\epsilon_0\) and \(\epsilon_1,\) we set \(\epsilon_0 = \frac{\epsilon}{m}\) and \(\epsilon_1 = \frac{\epsilon}{10m}.\)

Based on the triangle and 2-star counts, we estimated the clustering coefficient as \(\frac{3x f_\Delta(G)}{f_2(G)}\), where \(f_\Delta(G)\) (resp. \(f_2(G)\)) is the estimate of the triangle (resp. 2-star) count.

Figure 15 shows the relative errors of the triangle count, 2-star count, and clustering coefficient. Note that the relative error of the 2-star count is not changed by changing \(\mu^*\) because the 2-star algorithm does not use the ARR. Figure 15 shows that the relative error of the 2-star count is much smaller than that of the triangle count. This is because each user can count her 2-stars locally (whereas she cannot count her triangles), as described in Section 1. Consequently, the relative error of the clustering coefficient is almost the same as that of the triangle count, as the denominator \(f_2(G)\) in the clustering coefficient is very accurate.

Note that the clustering coefficient requires the privacy
baskets for calculating both \( \hat{f}_\Delta(G) \) and \( \hat{f}_{\Delta}(G) \) (in Figure 15, \( 2\varepsilon \) in total). However, we can accurately calculate \( \hat{f}_{\Delta}(G) \) with a very small privacy budget, as shown in Figure 15. Thus, we can accurately estimate the clustering coefficient with almost the same privacy budget as the triangle count by assigning a very small privacy budget (e.g., \( \varepsilon = 0.1 \) or \( 0.2 \)) for \( \hat{f}_{\Delta}(G) \).

In summary, we can accurately estimate the clustering coefficient as well as the triangle count under edge LDP by using our \( \text{ARROneNS}_\Delta \) with double clipping.

### D Experiments Using the Barabási-Albert Graph Datasets

In Section 6, we evaluated our algorithms using two real datasets. Below we also evaluate our algorithms using a synthetic graph based on the BA (Barabási-Albert) graph model [11], which has a power-law degree distribution.

In the BA graph model, a graph of \( n \) nodes is generated by attaching new nodes one by one. Each new node is connected to \( m \in \mathbb{Z}_{\geq 0} \) existing nodes, and each edge is connected to an existing node with probability proportional to its degree. We used NetworkX [28], a Python package for complex networks, to generate synthetic graphs based on the BA graph model.

We generated a graph \( G = (V, E) \) with the same number of nodes as Gplus; i.e., \( n = 107614 \) nodes. For the number \( m \) of edges per node, we set \( m = 50, 114, \) or \( 500 \). Using these graphs, we evaluated our three algorithms with double clipping. We set parameters in the same as Section 6; i.e., \( \alpha = 150, \beta = 10^{-6}, \varepsilon_0 = \frac{\varepsilon}{m}, \) and \( \varepsilon_1 = \varepsilon_2 = \frac{9\varepsilon}{m} \). For each algorithm, we averaged the relative error over 10 runs.

Figure 16 shows the results, where \( \varepsilon = 1 \) and \( \mu^* = 10^{-3} \). We observe \( \text{ARROneNS}_\Delta \) significantly outperforms \( \text{ARRFull}_\Delta \) and \( \text{ARRTwoNS}_\Delta \) when \( m = 500 \), and that \( \text{ARROneNS}_\Delta \) performs almost the same as \( \text{ARRFull}_\Delta \) when \( m = 50 \) or 114.

To examine the reason for this, we also decomposed the estimation error into two components (the first error by empirical estimation and the second error by the Laplacian noise) in the same way as Figure 11. Figure 17 shows the results. We also show in Table 3 the number \( C_4 \) of 4-cycles in each BA graph (\( m = 50, 114, \) or 500) and Gplus.

From Figure 17 and Table 3, we can explain Figure 16 as follows. The BA graphs with \( m = 50 \) and 114 have a much smaller number \( C_4 \) of 4-cycles than Gplus, as shown in Table 3. Consequently, the Laplacian noise is relatively large and dominant for these two graphs, as shown in Figure 17. In particular, the Laplacian noise is the largest in \( \text{ARRTwoNS}_\Delta \) because it cannot effectively reduce the global sensitivity by double clipping, as explained in Section 5. In contrast, the BA graph with \( m = 500 \) has a larger number \( C_4 \) of 4-cycles than Gplus, and therefore the Laplacian noise is not dominant (except for \( \text{ARRTwoNS}_\Delta \)). This explains the results in Figure 16.

![Figure 16: Relative error of our three algorithms with double clipping in the BA graphs (\( n = 107614, \varepsilon = 1, \mu^* = 10^{-3} \)).](image)

![Figure 17: Relative error of empirical estimation and the Laplacian noise in our three algorithms with double clipping in the BA graphs (\( n = 107614, \varepsilon = 1, \mu^* = 10^{-3} \)).](image)

| \( m \) | \( C_4 \) |
|---|---|
| 50 | \( 8.8 \times 10^8 \) |
| 114 | \( 1.7 \times 10^{10} \) |
| 500 | \( 3.1 \times 10^{12} \) |
| Gplus | \( 2.8 \times 10^{12} \) |

These results show that \( \text{ARROneNS}_\Delta \) outperforms \( \text{ARRFull}_\Delta \) especially when the number \( C_4 \) of 4-cycles is large. As we have shown in Section 6 and Appendix D, \( C_4 \) is large in a large graph (e.g., \( n \approx 10^6 \)) or dense graph (e.g., Gplus, BA graph with \( m = 500 \)).

### E Edge Clipping and Noisy Triangle Clipping

In Section 6, we showed that our double clipping significantly reduces the estimation error. To investigate the effect of edge clipping and noisy triangle clipping independently, we also performed the following ablation study.

We evaluated our three algorithms with only edge clipping; i.e., each user calculates a noisy degree \( \hat{d}_i \) (possibly with edge clipping) and then adds \( \text{Lap}(\frac{\hat{d}_i}{\sqrt{m}}) \) to her noisy triangle count. Then we compared them with our algorithms with double clipping and without clipping.

Figure 18 shows the results, where \( \varepsilon = 1, \mu^* = 10^{-6} \) or \( 10^{-3} \), and “EC” represents our algorithms with only edge clipping. We observe that “EC” outperforms “\( d_{\text{max}} \)” (w/o clipping) and is outperformed by “DC” (double clipping). The difference between “EC” and “DC” is significant especially when \( \mu^* = 10^{-6} \) (“DC” is smaller than \( \frac{1}{100} \) of “EC”). This is
Figure 18: Relative error of our three algorithms without clipping (“dmax”), with only edge clipping (“EC”), and with double clipping (“DC”) when \( \epsilon = 1 \) and \( \mu^* = 10^{-6} \) or \( 10^{-3} \) (\( n = 107614 \) in Gplus, \( n = 896308 \) in IMDB).

because our noisy triangle clipping reduces the global sensitivity by using a small value of \( \mu^* \). From Figure 18, we conclude that each component (i.e., edge clipping, noisy triangle clipping) is essential in our double clipping.

\section{Proof of Proposition 2}

Let \( A_\epsilon(A_i) = (A_\epsilon^1(A_i), A_\epsilon^2(M_i)(A_i)) \) be the randomized user by user \( i \), in the composition. To establish that \( A_\epsilon(A_i) \) satisfies \( \epsilon \)-edge LDP for every \( v_i \in V \), we will prove that (1) holds for \( A_\epsilon(A_i) \). To do this, first write

\[
\text{Pr}[A_\epsilon^1(A_i) = r_i] \text{Pr}[A_\epsilon^2(M_i)(A_i) = r_i^2 | A_\epsilon^1(A_i) = r_i^1] \leq e^{\epsilon^2} \text{Pr}[A_\epsilon^1(A_i) = r_i^1] \text{Pr}[A_\epsilon^2(M_i)(A_i) = r_i^2 | A_\epsilon^1(A_i) = r_i^1] = e^{\epsilon^1 + \epsilon^2} \text{Pr}[A_\epsilon^1(A_i) = r_i^1, A_\epsilon^2(M_i)(A_i) = r_i^2].
\]

This establishes the result. \( \Box \)

\section{Proof of Statements in Section 4}

\subsection{Proof of Theorem 1}

Let \( a_i, a_i' \in \{0,1\}^n \) be two neighbor lists that differ in one bit. Let \( t_i', s_i' \), and \( w_i' \) respectively be the values of \( t_i \) (line 1 of Algorithm 1), \( s_i \) (line 12), and \( w_i \) (line 13) when the neighbor list of user \( v_i \) is \( a_i' \). Let \( \Delta w_i = |w_i' - w_i| \). If we have \( t_i' - t_i \in [0,d_{\text{max}}] \) and \( s_i' - s_i \in [0,d_{\text{max}}] \), and therefore \( \Delta w_i = |t_i' - t_i| = \mu' (s_i' - s_i) \leq d_{\text{max}} \).

Since we add Lap \( \left( \frac{d_{\text{max}}}{\epsilon^2} \right) \) to \( w_i \), the second round provides \( \epsilon_2 \)-edge LDP. The first round uses \( ARR_{\epsilon_1, \mu} \) and provides \( \epsilon_1 \)-edge LDP. Thus, by sequential composition (Proposition 2), Algorithm 1 provides \( (\epsilon_1 + \epsilon_2) \)-edge LDP in total. It also provides \( (\epsilon_1 + \epsilon_2) \)-relationship DP because it uses only the lower-triangular part of \( A \) (Proposition 1).

\subsection{Proof of Theorem 2}

Unbiased Estimators. First, we will show that \( \hat{f}_\Delta(G) \) satisfies \( \mathbb{E}[\hat{f}_\Delta(G)] = f_\Delta(G) \) for all \( G \in \mathcal{G} \), in \( ARR_{\Delta_{\text{full}}} \), \( ARR_{\text{OneNS}_\Delta} \), \( ARR_{\text{TwoNS}_\Delta} \). Regardless of algorithm, we have

\[
\mathbb{E}[\hat{f}_\Delta(G)] = \frac{1}{\mu^*(1 - \rho)} \sum_{i = 1}^{n} \mathbb{E}[w_i] = \frac{1}{\mu^*(1 - \rho)} \sum_{i = 1}^{n} \mathbb{E}[\mu^* \rho s_i] = \frac{1}{\mu^*(1 - \rho)} \sum_{i = 1}^{n} \sum_{a_i = 1}^{\epsilon} \sum_{a_i = 1}^{\epsilon} \mathbb{E}[1_{(v_i, v_k) \in M_i} - \mu^* \rho],
\]

where \( \rho = e^{-\epsilon_1} \) and the quantities \( \mu^*, t_i, s_i \) are defined in Algorithm 1. Given that \( a_i = a_i = 1 \), we have that \( \Pr[(v_i, v_j) \in E'] = \Pr[(v_i, v_k) \in E'] = \mu \) by definition of ARR. Furthermore, \( \Pr[(v_j, v_k) \in E'] = \mu \) if \( a_j = a_k = 1 \), and \( \Pr[(v_j, v_k) \in E'] = 0 \) otherwise. Examining (7), (8), and (9), we have

\[
\Pr[(v_j, v_k) \in M_i] = \begin{cases} 
\mu^* & a_{j,k} = 0 \\
\mu^* \rho & a_{j,k} = 1 
\end{cases}
\]

for all the three algorithms (note that \( \mu^* = \mu^* \), \( \mu^2 \), and \( \mu^3 \) in \( ARR_{\Delta_{\text{full}}} \), \( ARR_{\text{OneNS}_\Delta} \), \( ARR_{\text{TwoNS}_\Delta} \), respectively). Thus, \( \mathbb{E}[1_{(v_i, v_k) \in M_i}] = \mu^* (\rho + (1 - \rho) a_{j,k}) = \mu^* \) if \( a_i = 1 \) and \( \mu^* \rho \) if \( a_i = 0 \). Plugging into (19), we have

\[
\mathbb{E}[\hat{f}_\Delta(G)] = \frac{1}{\mu^*(1 - \rho)} \sum_{i = 1}^{n} \sum_{a_i = a_i = 1}^{\epsilon} \mu^*(\rho + (1 - \rho) a_{j,k}) - \mu^* \rho = \frac{1}{\mu^*(1 - \rho)} \sum_{i = 1}^{n} \sum_{a_i = a_i = 1}^{\epsilon} \mu^*(1 - \rho) a_{j,k} = \sum_{i = 1}^{n} \sum_{a_i = a_i = 1}^{\epsilon} a_{j,k} = f_\Delta(G).
\]

Thus, \( \hat{f}_\Delta(G) \) is unbiased. \( \Box \)
**$l_2$ Loss of Estimators.** Using bias-variance decomposition, we have for any graph $G$,

$$
I_2^2(\hat{f}(G), \hat{f}(G)) = \mathbb{E}[(\hat{f}(G) - f(G))^2] = \mathbb{E}[(\hat{f}(G) - \mathbb{E}[\hat{f}(G)])^2] + \mathbb{V}ar[\hat{f}(G)] = \mathbb{V}ar[\hat{f}(G)],
$$

where the last step follows because $\hat{f}$ is unbiased. Since $\hat{f}(G) = \frac{1}{\mu(1-\rho)} \sum_{i=1}^{n} \hat{w}_i$, we have

$$
\mathbb{V}ar[\hat{f}(G)] = \frac{1}{(\mu')^2(1-\rho)^2} \mathbb{V}ar \left[ \sum_{i=1}^{n} \hat{w}_i \right] = \frac{1}{(\mu')^2(1-\rho)^2} \mathbb{V}ar \left[ \sum_{i=1}^{n} w_i + \text{Lap}(\frac{d_{\text{max}}}{\varepsilon_2}) \right] = \frac{1}{(\mu')^2(1-\rho)^2} \left( \mathbb{V}ar \left[ \sum_{i=1}^{n} w_i \right] + n \mathbb{V}ar \left[ \text{Lap}(\frac{d_{\text{max}}}{\varepsilon_2}) \right] \right) = \frac{1}{(\mu')^2(1-\rho)^2} \left( \mathbb{V}ar \left[ \sum_{i=1}^{n} w_i \right] + 2n \frac{d_{\text{max}}^2}{\varepsilon_2^2} \right), \tag{20}
$$

where the fourth line follows from independence of the added of Laplace noise. Now, we will prove bounds on $\mathbb{V}ar[\sum_{i=1}^{n} w_i]$ for $ARRFull_\Delta$, $ARROneNS_\Delta$, and $ARRTwoNS_\Delta$. In the following, we let $S_k(G)$ be the number of $k$-stars in $G$ and $C_4(G)$ be the number of 4-cycles in $G$.

**Bounding the Variance in $ARRFull_\Delta$.** In $ARRFull_\Delta$, $M_i$ is defined by (7). Thus we have

$$
\sum_{i=1}^{n} w_i = \sum_{i=1}^{n} \sum_{1 \leq j < k \leq i \leq n}^{\text{any order}} \frac{1}{a_{i,j} a_{i,k}} \mathbb{1}_{(v_j, v_k) \in E'},
$$

For $j < k$, we can introduce the constant $c_{jk} = \sum_{k' \leq n} a_{ij} a_{ik}$. Notice that for any two nodes $j$ and $k$, $\mathbb{1}_{(v_j, v_k) \in E'}$ for $1 \leq j \leq k$ are mutually independent, because all edges in $E'$ are mutually independent. Furthermore, the indicator $\mathbb{1}_{(v_j, v_k) \in E'}$ is a Bernoulli random variable with parameter either $\mu$ or $\mu \rho$, and in either case, $\mathbb{V}ar[\mathbb{1}_{(v_j, v_k) \in E'}] \leq \mu$. We have

$$
\mathbb{V}ar \left[ \sum_{i=1}^{n} w_i \right] = \sum_{1 \leq j < k \leq n} \mathbb{V}ar \left[ \sum_{i=1}^{n} \mathbb{1}_{(v_j, v_k) \in E'} c_{jk} \right] = \sum_{1 \leq j < k \leq n} \mathbb{V}ar \left[ \sum_{i=1}^{n} \mathbb{1}_{(v_j, v_k) \in E'} c_{jk} \right] = \sum_{1 \leq j < k \leq n} \frac{\mu^2 c_{jk}^2}{\mu^2},
$$

By Lemma 1 (which is shown at the end of Appendix G.2), we have $\sum_{1 \leq j < k \leq n} c_{jk}^2 \leq 2C_4(G) + S_2(G)$. Plugging into (20) (and substituting $\rho' = \mu$), we obtain

$$
\mathbb{V}ar[\hat{f}(G)] = \frac{1}{(1-\rho)^2} \left( \frac{1}{\mu} (2C_4(G) + S_2(G)) + 2n \frac{d_{\text{max}}^2}{\varepsilon_2^2 \mu^2} \right).
$$

This establishes the result. $\square$

**Bounding the Variance in $ARROneNS_\Delta$.** In $ARROneNS_\Delta$, for a fixed $v_i \in V$, we have $(v_j, v_k) \in E$, and $(v_i, v_k) \in E'$ from (8), thus,

$$
\sum_{1 \leq j < k \leq n} \mathbb{1}_{(v_j, v_k) \in E'} = \sum_{1 \leq j < k \leq n} a_{ij} a_{ik} \mathbb{1}_{(v_j, v_k) \in E'}.
$$

Define the random variable $F_{ijk} = \mathbb{1}_{(v_j, v_k) \in E'} \mathbb{1}_{(v_i, v_k) \in E'}$. Substituting, we have

$$
\sum_{i=1}^{n} w_i = \sum_{1 \leq j < k \leq n} a_{ij} a_{ik} F_{ijk}.
$$

The set $\{i, j', j, k' \}$ takes five or more distinct values, then $F_{ijk}$ and $F_{ij'k'}$ involve distinct edges and are independent random variables. Thus, $\text{Cov}(F_{ijk}, F_{ij'k'}) = 0$. Otherwise, the events are not independent, and we will use the upper bound $\text{Cov}(F_{ijk}, F_{ij'k'}) \leq \mathbb{E}[F_{ijk} F_{ij'k'}] = \text{Pr}[F_{ijk} = F_{ij'k'} = 1]$, which holds because the $F_{ijk}$ have domain $\{0, 1\}$. Thus

$$
\mathbb{V}ar \left[ \sum_{i=1}^{n} w_i \right] \leq \sum_{1 \leq j < k \leq n} \sum_{1 \leq j < k' < \leq n} a_{ij} a_{ik} a_{ij'} a_{ik'} \text{Pr}[F_{ijk} = F_{ij'k'} = 1].
$$

Define a choice of $(i, j, k, i', j', k') \in [n]^6$ to be a valid choice if $j < k < i$, $j' < k' < i'$ (ordering requirement), $a_{ijk} = a_{ij'} = a_{ik'} = 1$ (edge requirement), and $3 \leq \{i, j, j', k, k'\} \leq 4$ (size requirement). We can write the above sum as

$$
\mathbb{V}ar \left[ \sum_{i=1}^{n} w_i \right] \leq \sum_{i, j, k, i', j', k', k'} \text{Pr}[F_{ijk} = F_{ij'k'} = 1].
$$

In the above sum, each valid choice implies there exists a subgraph of $G$ that associates each of $\{v_i, v_j, v_j, v_k, v_k'\}$...
with one node in the subgraph and contains edges $(v_i, v_j), (v_i, v_k), (v_j, v_p), (v_j, v_k)$. Conversely, each subgraph of $G$ of three or four nodes can have a certain number of valid choices mapped to it. For each subgraph, we now go over the number of possible valid choices that can map to it:

1. 4-cycle: By ordering, either $i$ or $i’$ is mapped to the node of the 4-cycle with maximal index. WLOG, suppose $i$ is mapped to this node. By edge requirements, $i’$ has an index equal to the opposite node in the 4-cycle. By ordering, there is now one way to map $j, j’, k, k’$. Thus, each 4-cycle can be associated with 2 valid choices.

2. 3-path: Consider the middle node $v_l$ in the 3-path (path graph on 4 nodes) that has the second-largest index. By ordering, either $i = l$ or $i’ = l$. WLOG, suppose $i = l$. Then, by the edge requirement $a_{ij, j'} = a_{ij, k} = 1$, the middle node other than $v_l$ is $j’$. However, this means either $i = j’$ or $i = k’$, and we have $j’ > i’$ or $k’ > i’$. This violates the order requirement, and therefore there are 0 valid choices.

3. 3-star: By edge requirement, both $i, i’$ map to the central node in the 3-star. $j, j’, k, k’$ can map to the other three nodes in any way that satisfies ordering. Suppose the three nodes are $v_a, v_b, v_c$ with $a < b < c$. Only one of the three nodes can be duplicated in this mapping. For example, if $a$ is duplicated, then both $j$ and $j’$ map to $v_a$, and there are two remaining choices for how to map $k$ and $k’$. Thus, each $S_3$ can be associated with 6 valid choices.

4. Triangle: Either $i$ or $i’$ maps to the maximal node in the triangle by ordering. WLOG, suppose $i$ does. By the edge requirement, $i’$ maps to a different node. However, this means $i = k’$ or $i = j’$, so $k’ > i’$, contradicting ordering. Thus, there are 0 valid choices.

5. 2-star: By edge requirements, both $i$ and $i’$ map to the central node in the 2-star. We then have just one mapping for the remaining indices. Thus, there is 1 valid choice.

Figure 19 shows an example of two 4-cycles, six 3-stars, and one 2-stars.

Putting this together,

$$\mathbb{V} \left[ \sum_{i=1}^{n} w_i \right] \leq 2C_4(G)\mu^3 + 6S_3(G)\mu^3 + S_2(G)\mu^2.$$

Plugging into (20), we get

$$\mathbb{V}[\hat{f}_{\Delta}(G)]$$

$$\leq \frac{1}{(1 - \rho)^2} \left( \frac{1}{\mu^2} (2C_4(G) + 6S_3(G)) + \frac{\mu}{\mu^2} S_2(G) + 2n^2 d_{max}^2/\mu^3 E_2^2 \right).$$

This establishes the result.

**Bounding the Variance in ARRTwoNS$_\Delta$.** In ARRTwoNS$_\Delta$, for a fixed $v_i \in V$, we have $(v_j, v_k) \in E_i$ if and only if $j < k < i$. $(v_j, v_k) \in E'$, $(v_i, v_k) \in E'$, and $(v_i, v_k) \in E'$ from (9). Thus,

$$\sum_{i=1}^{n} w_i = \sum_{i=1}^{n} \sum_{1 \leq j < k < n} \sum_{a_{ij} = 1} a_{ij} a_{ik} a_{i j k} (v_j, v_k) \in E_i, \sum_{1 \leq j < k < n} a_{ij} a_{ik} a_{ij k} (v_j, v_k) \in E_i, \sum_{1 \leq j < k < n} a_{ij} a_{ik} a_{ij k} (v_j, v_k) \in E_i.$$

Define the random variable $F_{ijk} = I_{(v_j, v_k) \in E_i} I_{(v_i, v_k) \in E'}. I_{(v_i, v_k) \in E'}. I_{(v_i, v_k) \in E'}. I_{(v_i, v_k) \in E'}$. Following the same steps as those in the proof of ARROneNS$_\Delta$, we have

$$\mathbb{V} \left[ \sum_{i=1}^{n} w_i \right] = \sum_{1 \leq j < k < n} a_{ij} a_{ik} a_{j k} a_{j k} \text{ Cov}(F_{ijk}, F_{j’k’})$$

$$\leq \sum_{i, j, j’, k, k’} \text{ Pr}[F_{ijk} = F_{j’k’}] = 1.$$

As we showed in the proof for ARROneNS$_\Delta$, each 4-cycle of $G$ has at most 2 valid choices mapped to it, each 3-star of $G$ has at most 6 valid choices mapped to it, and each 2-star of $G$ has at most one valid choice mapped to it.

In the following, let $P(i, j, k)$ be the event that $F_{ijk} = F_{j’k’} = 1$. Observing Figure 4, we can see that in both possible ways in which a valid choice maps to a 4-cycle, then $P(i, j, k)$ holds when at least 3 edges in $E'$ are present. Each edge in $E'$ is independent, and is present with probability at most $\mu$. Thus, $\text{Pr}[P(i, j, k)] \leq \mu^3$. Next, if a valid choice maps to a 3-star, then $P(i, j, k)$ implies at least 3 edges in $E'$ are present. Thus, $\text{Pr}[P(i, j, k)] = 1 \leq \mu^3$. Finally, if a valid choice maps to a 2-star, then $P(i, j, k) = 1$ if and only if 2 edges in $E'$ are present. Thus, $\text{Pr}[P(i, j, k) = 1] \leq \mu^2$. As we showed in the proof for ARROneNS$_\Delta$, each 4-cycle of $G$ has at most 2 valid choices mapped to it, each 3-star of $G$ has at most 6 valid choices mapped to it, and each 2-star of $G$ has at most one valid choice mapped to it.

In the following, let $P(i, j, k)$ be the event that $F_{ijk} = F_{j’k’} = 1$. Observing Figure 4, we can see that for each possible mapping of a valid choice to a 4-cycle, five edges must be present in $G'$ in order for $P(i, j, k) = 1$. Thus, $\text{Pr}[P(i, j, k) = 1] \leq \mu^4$. For each possible mapping of a valid choice to a 3-star, five edges must be present in $G'$ in order for $P(i, j, k) = 1$. Thus, $\text{Pr}[P(i, j, k)] \leq \mu^2$. As we showed in the proof for ARROneNS$_\Delta$, each 4-cycle of $G$ has at most 2 valid choices mapped to it, each 3-star of $G$ has at most 6 valid choices mapped to it, and each 2-star of $G$ has at most one valid choice mapped to it.
Thus, \( \Pr[P(i, j, k) = 1] \leq \mu^3 \). For each possible mapping of a valid choice to a 2-star, three edges must be present in \( G' \) in order for \( P(i, j, k) = 1 \). Thus, \( \Pr[P(i, j, k) = 1] \leq \mu^3 \).

Plugging into (20), we get
\[
\forall \left[ f \right] (G) \leq \frac{1}{(1 - p)^2} \left( \frac{1}{\mu} 2C_4(G) + 6S_3(G) + \frac{1}{\mu^3} S_2(G) + 2n \frac{d_{\max}^2}{\mu^2 \epsilon_n^2} \right)
\]
This establishes the result. \( \square \)

**Lemma 1.** Let \( c_{ij} = \sum_{i < j} a_{ij}a_{ij} \). Then,
\[
\sum_{i < j} c_{ij}^2 \leq 2C_4(G) + S_2(G).
\]

**Proof.**
\[
\sum_{i < j} c_{ij}^2 = \sum_{i < j} c_{ij} + \sum_{i < j} c_{ij}(c_{ij} - 1) = S_2(G) + \sum_{i < j} c_{ij}(c_{ij} - 1).
\]

Let \( C_{i-1-\ldots-1} \) be the number of 4-cycles in \( G \) such that the first and third nodes are \( v_i \) and \( v_j \), respectively \((i < j)\), and the remaining two nodes have smaller indices than \( i \). From middle nodes in 2-paths starting at \( v_i \) and ending at \( v_j \), we can choose two nodes as the second and fourth nodes in the 4-cycles. \( c_{ij} \) is the number of nodes that have smaller IDs than \( v_i \) and are connected to \( v_i \) and \( v_j \). Thus, \( C_{i-1-\ldots-1} \) is equal to the probability that the number of successes in the binomial distribution \( \binom{n}{2} \). Therefore, we have
\[
\sum_{i < j} c_{ij} = S_2(G) + \sum_{i < j} 2C_{i-1-\ldots-1} \leq S_2(G) + 2C_4(G).
\]

The last inequality comes from the fact that two nodes with the largest indices may not be opposite to each other in some 4-cycles in \( G \). \( \square \)

**H.1 Proof of Statements in Section 5**

**H.2 Proof of Theorem 4**

Recall that \( t_{i,j} = \| \{v_i, v_j, v_k\} : a_{ijk} = 1, (v_i, v_k) \in M_i, j < k < i \| \). Let \( t'_{i,j} = \| \{v_i, v_j, v_k\} : a_{ijk} = 1, (v_i, v_k) \in M_i \| \). Then \( t_{i,j} \leq t'_{i,j} \). Thus we have
\[
\Pr(t_{i,j} > \kappa_1) \leq \Pr(t_{i,j} > \kappa_1) \leq \Pr(t'_{i,j} \geq \kappa_1).
\]

Below we first prove (12) and (13). Then we prove (14).

**Proof of (12) and (13).** For each edge \((v_i, v_j)\), we have \( \sum_{a \in V} 1_{(v_i, v_j) \in a} \leq \bar{d}_i \). In ARRFull, each edge \((v_i, v_j)\) is included in \( E' \) with probability at most \( \mu \), and all the events are independent. In ARROneNS, each of the edges \((v_i, v_j)\) and \((v_i, v_k)\) is included in \( E' \) with probability at most \( \mu \), and all the events are independent. Thus, \( \Pr(t'_{i,j} \geq \kappa_1) \) is less than or equal to the probability that the number of successes in the binomial distribution \( B(\bar{d}_i, \mu^*) = B(\bar{d}_i, \mu^*) \) (\( \mu^* = \mu \) in ARRFull and \( \mu^2 \) in ARROneNS) is larger than or equal to \( \kappa_1 \).

Let \( X_{n,p} \) be a random variable representing the number of successes in the binomial distribution \( B(n, p) \). For \( \kappa_1 \geq \mu_{\kappa_1} \), we have
\[
\Pr(t'_{i,j} \geq \kappa_1) \leq \Pr(X_{n,p} \geq \kappa_1) = F(\kappa_1 ; n, p) = \Pr(\kappa_1 ; n, p) \geq \kappa_1) \quad \text{i.e., } F \text{ is a cumulative distribution function of } B(n, p).
\]

Since \( \kappa_1 \geq \mu_{\kappa_1} \), we have
\[
\Pr(t'_{i,j} \geq \kappa_1) \leq \Pr(t'_{i,j} \geq \kappa_1) \leq \Pr(t'_{i,j} \geq \kappa_1).
\]
**Proof of (14).** Assume that $\kappa_i \geq \mu^2 \hat{d}_i$ in ARRTwoNS$_\Delta$. For each edge $(v_i, v_j)$, we have $\sum_{k \neq i,j} 1_{(i,k) \in E} \leq \hat{d}_i$. In addition, each of the edges $(v_i, v_k)$ and $(v_j, v_k)$ are included in $E'$ with probability at most $\mu$, and all the events are independent.

If $(v_i, v_j)$ is included in $E'$ (which happens with probability at most $\mu$), $Pr(t'_{i,j} \geq \kappa_i)$ is less than or equal to the probability that the number of successes in the binomial distribution $B(\hat{d}_i, \mu^2)$ is larger than or equal to $\kappa_i$. Otherwise (i.e., if $(v_i, v_j)$ is not included in $E'$), then $t'_{i,j} = 0$.

Thus, if $\kappa_i \geq \mu^2 \hat{d}_i$, we have

$$Pr(t'_{i,j} \geq \kappa_i) \leq \mu Pr(X_d, \mu^2 \geq \kappa_i) = \mu F(\hat{d}_i - \kappa_i; \hat{d}_i, 1 - \mu^2) \leq \mu \exp \left[ -\hat{d}_i D \left( \frac{\hat{d}_i - \kappa_i}{\hat{d}_i} \parallel 1 - \mu^2 \right) \right] \text{(by Chernoff bound)}$$

and therefore (14) holds (as $Pr(t_{i,j} > \kappa_i) \leq Pr(t'_{i,j} \geq \kappa_i)$).

If $\mu^2 \hat{d}_i \leq \kappa_i < \mu^2 \hat{d}_i$, (14) can be written as: $Pr(t_{i,j} > \kappa_i) \leq \mu \exp \left[ -\hat{d}_i D \left( \frac{\mu^2}{\hat{d}_i} \parallel \frac{\mu^2}{\kappa_i} \right) \right] = \mu$. This clearly holds because each edge $(v_i, v_j)$ is included in $E'$ with probability at most $\mu$. \qed

I Addendum to Double Clipping

I.1 Sensitivity

After this paper was published at USENIX Security 2022\(^3\), we found that we underestimated the sensitivity of our double clipping technique in Section 5. For example, consider a graph $G = (V, E)$, where $V = \{v_1, v_2, v_3, v_4\}$ and $E = \{(v_1, v_4), (v_2, v_3), (v_3, v_4)\}$. Assume that ARRFull$_\Delta$ with double clipping is used as a triangle counting algorithm and that a message $M_4$ for $v_4$ is $M_4 = \{(v_1, v_3), (v_2, v_3)\}$. Then, $G$ has two noisy triangles $(v_1, v_3, v_4)$ and $(v_2, v_3, v_4)$ whose noisy edges are $(v_1, v_3)$ and $(v_2, v_3)$, respectively. Assume that $\kappa$ is set to $\kappa_4 = 1$.

Recall that for $j$ such that $a_{i,j} = 1$ and $j < i$, we calculate $t_{i,j} = \min \{(v_i, v_j, v_k) : a_{i,k} = 1, (v_j, v_k) \in E, j < k < i\}$ (see line 5 in Algorithm 2). In the above example, we have that $t_{i,4} = t_{4,2} = 1$ and $t_{4,3} = 0$. Thus, no triangle clipping occurs for $v_4$, and consequently, $t_i = 2$ (see line 8 in Algorithm 2).

Now, consider a graph $G' = (V, E')$ obtained by removing $(v_3, v_4)$ from $G$, i.e., $E' = \{(v_1, v_4), (v_2, v_4)\}$. $G'$ has no noisy triangles, as $(v_3, v_4)$ is removed from two noisy triangles $(v_1, v_3, v_4)$ and $(v_2, v_3, v_4)$ in $G$. Thus, $t_{i,4} = t_{4,2} = t_{4,3} = 0$, and consequently, $t_i = 0$. This means that the sensitivity of $t_i$ is 2. Therefore, $\kappa_i (= 1)$ is not an upper bound on the sensitivity of $t_i$. The root cause of this issue is that adding a single edge of user $v_i$ could increment $t_{i,j}$ for more than $\kappa_i$ choices of $v_j$ (2 choices in the above example), even if each $t_{i,j}$ is clipped to be at most $\kappa_i$.\footnote{1https://www.usenix.org/system/files/sec22-imola.pdf}

I.2 $(\epsilon, \delta)$-edge LDP

The sensitivity issue in Appendix I.1 implies that adding Lap($\frac{\kappa_i}{\delta^2}$) (line 11 in Algorithm 2) does not provide $\epsilon_2$-edge LDP. Consequently, our algorithms with double clipping do not satisfy $(\epsilon_0 + \epsilon_1 + \epsilon_2)$-edge LDP; i.e., Theorem 3 is incorrect.

However, our algorithms with double clipping provide $(\epsilon, \delta)$-edge LDP and $(\epsilon, \delta)$-relationship DP, where $\epsilon = \epsilon_0 + \epsilon_1 + \epsilon_2$ and $\delta$ is immediately derived from Theorem 4 that upper bounds the triangle excess probability. Fortunately, our double clipping is very effective even if $\delta$ is extremely small (e.g., $\delta \ll n^{-1}$ or $n^{-2}$), as shown in Appendix I.3.

Formally, we define $(\epsilon, \delta)$-edge LDP and $(\epsilon, \delta)$-relationship DP as follows:

**Definition 4** $(\epsilon, \delta)$-edge LDP. Let $\epsilon \in \mathbb{R}_{>0}$ and $\delta \in [0, 1]$. For $i \in [n]$, let $R_i$ be a local randomizer of user $v_i$ that takes $a_i$ as input. We say $R_i$ provides $(\epsilon, \delta)$-edge LDP if for any two neighbor lists $a_i, a'_i \in \{0, 1\}^n$ that differ in one bit and any $S \subseteq \text{Range}(R_i)$,

$$Pr[R_i(a_i) \in S] \leq e^\epsilon Pr[R_i(a'_i) \in S] + \delta.$$ 

**Definition 5** $(\epsilon, \delta)$-relationship DP. Let $\epsilon \in \mathbb{R}_{>0}$ and $\delta \in [0, 1]$. For $i \in [n]$, let $R_i$ be a local randomizer of user $v_i$ that takes $a_i$ as input. We say $R_i$ provides $(\epsilon, \delta)$-relationship DP if for any two neighboring graphs $G, G' \in G$ that differ in one edge and any $S \subseteq \text{Range}(R_i) \times \ldots \times \text{Range}(R_n)$,

$$Pr[(R_1(a_1), \ldots, R_n(a_n)) \in S] \leq e^\epsilon Pr[(R_1(a'_1), \ldots, R_n(a'_n)) \in S] + \delta,$$

where $a_i$ (resp. $a'_i$) $\in \{0, 1\}^n$ is the $i$-th row of the adjacency matrix of graph $G$ (resp. $G'$).

**Proposition 3** $(\epsilon, \delta)$-edge LDP and $(\epsilon, \delta)$-relationship DP. If each of local randomizers $R_1, \ldots, R_n$ provides $(\epsilon, \delta)$-edge LDP, then $(R_1, \ldots, R_n)$ provides $(2\epsilon, 2\delta)$-relationship DP. Additionally, if each $R_i$ uses only bits $a_i, \ldots, a_{i-1}$ for users with smaller IDs (i.e., only the lower triangular part of $A$), then $(R_1, \ldots, R_n)$ provides $(\epsilon, \delta)$-relationship DP.

**Proof.** Immediately derived from group privacy [25]. \qed

Then, our algorithms with double clipping provide the following privacy guarantees:

**Theorem 5.** For $i \in [n]$, let $R_i^1, R_i^2$ be the randomizers used by user $v_i$ in rounds 1 and 2 of our algorithms with double clipping (Algorithms 1 and 2). Assume that the clipping
threshold $\kappa_i$ is set so that the upper-bound in Theorem 4 is smaller than or equal to $\beta \in [0,1]$; i.e.,
\[
\exp\left(-\tilde{d}_i D\left(\frac{\kappa_i}{d_i} \mid \mu\right)\right) \leq \beta \tag{21}
\]
\[
\exp\left(-\tilde{d}_i D\left(\frac{\kappa_i}{d_i} \mid \mu^2\right)\right) \leq \beta \tag{22}
\]
\[
\mu \exp\left[-\tilde{d}_i D\left(\frac{\max(\kappa_i, \mu d_i)}{d_i} \mid \mu^2\right)\right] \leq \beta \tag{23}
\]
in AARRFull, AARROneNS, and AARRTwoNS, respectively. Let $\mathcal{R}_t(a_i) = (\mathcal{R}_1^t(a_i), \mathcal{R}_2^t(M_t)(a_i))$ be the composition of the two randomizers. Then, $\mathcal{R}_t$ satisfies $(\varepsilon, \delta)$-edge LDP, and $(\mathcal{R}_1, \ldots, \mathcal{R}_n)$ satisfies $(\varepsilon, \delta)$-relationship DP, where $\varepsilon = \varepsilon_0 + \varepsilon_1 + \varepsilon_2$ and $\delta = n\beta$.

**Proof.** We first show that $\mathcal{R}_t$ satisfies $(\varepsilon, \delta)$-edge LDP, where $\varepsilon = \varepsilon_0 + \varepsilon_1 + \varepsilon_2$. For $S \subseteq \text{Range}(\mathcal{R}_t)$, let $S_0 \subseteq S$ be the set such that
\[
\Pr[\mathcal{R}_t(a_i) = s] \leq e^\varepsilon \Pr[\mathcal{R}_t(a_i') = s] \tag{24}
\]
holds for any $s \in S_0$ and
\[
\Pr[\mathcal{R}_t(a_i) = s] > e^\varepsilon \Pr[\mathcal{R}_t(a_i') = s]
\]
holds for any $s \in S \setminus S_0$. Then, we have
\[
\Pr[\mathcal{R}_t(a_i) \in S] = \Pr[\mathcal{R}_t(a_i) \in S_0] + \Pr[\mathcal{R}_t(a_i) \in S \setminus S_0] \\
\leq e^\varepsilon \Pr[\mathcal{R}_t(a_i') \in S_0] + \delta,
\]
where
\[
\delta = \Pr[\mathcal{R}_t(a_i) \in S \setminus S_0].
\]
By (21), (22), (23), and Theorem 4, we always have
\[
\Pr(t_{i,j} > \kappa_i) \leq \beta.
\]
When $t_{i,1}, \ldots, t_{i,i-1} \leq \kappa_i$, $\kappa_i$ upper bounds the sensitivity of $t_i$, hence (24) holds (see Appendix H.1 for the proof in this case). Thus, $\delta$ can be written as follows:
\[
\delta \leq 1 - \Pr(t_{i,1}, \ldots, t_{i,i-1} \leq \kappa_i) \leq \sum_{j=1}^{i-1} \Pr(t_{i,j} > \kappa_i) \leq n\beta.
\]
Therefore, $\mathcal{R}_t$ satisfies $(\varepsilon, \delta)$-edge LDP, where $\varepsilon = \varepsilon_0 + \varepsilon_1 + \varepsilon_2$ and $\delta = n\beta$.

Next, we show that $(\mathcal{R}_1, \ldots, \mathcal{R}_n)$ satisfies $(\varepsilon, \delta)$-relationship DP. This is immediately derived from Proposition 3 and the fact that each $\mathcal{R}_t$ uses only the lower triangular part of $A$.

**Remark.** In line 8 of Algorithm 2, we calculate the total number $t_i$ of noisy triangles for user $v_i$ by clipping the number of noisy triangles; i.e., $t_i = \sum_{j=1,j<i} \min\{t_{i,j}, \kappa_i\}$. Theorem 5 holds even when we do not clip the number of noisy triangles and calculate $t_i$ by simply summing $t_{i,j}$.

---

**I.3 Experimental Results**

We evaluated the relationship between the relative error and $\delta (= n\beta)$ in our three algorithms with double clipping (“DC”) $(\varepsilon = 1, \mu^t = 10^{-3})$, $\beta = 10^{-6}, 10^{-12}, 10^{-18}, 10^{-24}, 10^{-30}$, or $10^{-36}$, $n = 107614$ in Gplus, $n = 896308$ in IMDB). Note that $\delta = 0$ in our three algorithms without double clipping (“$d_{max}$”).

**Figure 20:** Relationship between the relative error and $\delta$ $(= n\beta)$ in our three algorithms with double clipping (“DC”) $(\varepsilon = 1, \mu^t = 10^{-3})$. $\beta = 10^{-6}, 10^{-12}, 10^{-18}, 10^{-24}, 10^{-30}, 10^{-36}$, $n = 107614$ in Gplus, $n = 896308$ in IMDB). Note that $\delta = 0$ in our three algorithms without double clipping (“$d_{max}$”).

---

i.e., $t_i = \sum_{a,j=1,j<i} t_{i,j}$. In this case, we perform calculate $\tilde{d}_i$ by edge clipping and then set $\kappa_i$ so that the triangle excess probability is smaller than or equal to $\beta$; i.e., (21), (22), and (23) hold. Then, the failure probability for user $v_i$ is upper bounded by $\delta = n\beta$, and consequently, our algorithms provide $(\varepsilon, \delta)$-edge LDP and $(\varepsilon, \delta)$-relationship DP. When $\delta = n\beta$ is extremely small (e.g., $\delta \ll n^{-1}$ or $n^{-2}$), triangle removal rarely occurs. Thus, whether or not we clip the number of noisy triangles in line 8 of Algorithm 2 makes little difference to the utility. Therefore, in Appendix I.3, we evaluate our algorithms without modifying line 8 of Algorithm 2.

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Figure 21: Relative error of our three algorithms with (“DC”) or without (“$d_{max}$”) double clipping when $\beta = 10^{-14}$ ($n = 107614$ and $\delta = 1.08 \times 10^{-9} \ll n^{-1}$ in Gplus; $n = 896308$ and $\delta = 8.96 \times 10^{-9} \ll n^{-1}$ in IMDB). $\text{ARRFull}_\alpha (d_{max})$ is the algorithm in [30]. $\text{Cost}_{DL}$ is an upper-bound in (10).

or less while keeping relative error $= 0.14$ (resp. 0.21). For IMDB, our $\text{ARROneNS}_\triangle (DC)$ with $\beta = 10^{-14}$ (resp. $10^{-24}$) reduces the communication cost from 400 Gbits (6 hours) to 1.6 Gbits (80 seconds) or less while keeping relative error $= 0.097$ (resp. 0.16).

In summary, our algorithms with double clipping provide $(\varepsilon, \delta)$-edge LDP and $(\varepsilon, \delta)$-relationship DP, and dramatically reduce the communication cost of the algorithm in [30] even when $\delta \ll n^{-1}$ or $n^{-2}$.

Figure 22: Relative error of our three algorithms with (“DC”) or without (“$d_{max}$”) double clipping when $\beta = 10^{-24}$ ($n = 107614$ and $\delta = 1.08 \times 10^{-19} \ll n^{-2}$ in Gplus; $n = 896308$ and $\delta = 8.96 \times 10^{-19} \ll n^{-2}$ in IMDB). $\text{ARRFull}_\alpha (d_{max})$ is the algorithm in [30]. $\text{Cost}_{DL}$ is an upper-bound in (10).