Clinical, neurocognitive and demographic factors associated with functional impairment in the Australian Brain and Mind Youth Cohort Study (2008–2016)

Rico S C Lee,1,2 Daniel F Hermens,¹ Sharon L Naismith,1,3 Manreena Kaur,1,4 Adam J Guastella,1 Nick Glozier,1,5 Jan Scott,6,7 Elizabeth M Scott,1 Ian B Hickie1

In recent decades, early intervention services for youth with emerging mental disorders have extended their targets beyond those at risk of psychosis to also encompass those presenting with mood as well as other developmental and anxiety disorders. This approach creates several significant challenges. For example, some youth with depressive and anxiety disorders will ultimately develop psychotic or bipolar disorders; likewise, only a proportion of those receiving a diagnosis of bipolar disorders will consistently receive this diagnosis over the following 10 years.1 The lack of diagnostic stability in help-seeking youth reflects the evolving disease process and means that the illness trajectory is less certain than for older adults with established illness.2 3 From a research perspective, the use of dimensional approaches to phenomenology has helped us to understand illness progression in these early clinical stages, while from a clinical perspective, care and treatment have increasingly considered transdiagnostic interventions addressing core factors that may influence prognosis irrespective of cross-sectional diagnosis (eg,
findings have been replicated in large cohort studies,\(^1\) the impact of affective and positive symptoms on functioning remains more equivocal. More recently, these symptoms in a one-size-fits-all approach is also emphasised by the WHO.\(^9\) The recognition that more personalised interventions are urgently needed to enhance functioning and quality of life rather than simply targeting diagnosis-specific symptoms in a one-size-fits-all approach is also emphasised by the WHO.\(^9\) Given this interest in enhancement of functioning across all stages of mental illness and for youth and adults presenting to mental health services, it is therefore useful to examine the role of other key clinical (eg, medication exposure) and demographic factors (eg, age, gender) in determining functioning which would contribute to prognosis and attempts at personalised medicine.

Most path modelling studies to date have used small, single-diagnosis or dual-diagnosis cohorts, predominantly in individuals with a chronic mental illness. Findings consistently demonstrate that neurocognition and negative symptoms are robust predictors of functional outcome in schizophrenia and bipolar disorder.\(^7\)\(^1\)\(^2\)\(^3\) By contrast, the impact of affective and positive symptoms on functioning remains more equivocal. More recently, these findings have been replicated in large cohort studies,\(^1\)\(^3\)\(^4\) although the vast majority of existing studies have focused exclusively on schizophrenia. There have been no well-powered studies examining a mental disorder other than schizophrenia, such as affective disorders, despite depression being the leading cause of disability worldwide.\(^5\)\(^6\) Studies have also largely sidestepped the issue of psychotropic medication use. Furthermore, given that more than 75% of mental illnesses emerge before the age of 25,\(^3\) examining younger cohorts is critically important for the development of novel approaches to early intervention since most studies to date have targeted older individuals.\(^1\)

In order to build on prior research, a transdiagnostic and dimensional approach is ideally positioned to disentangle the factors associated with functioning. Key to this research strategy is the examination of shared constructs (eg, neurocognition) with clear links to pathophysiology,\(^6\)\(^7\)\(^8\) which can inform novel therapeutics that target specific neural circuitries.\(^1\)\(^7\)\(^8\) Transdiagnostic studies are also able to harness the variance across disorders, with the goal of developing robust, unifying models that are explanatory in nature.\(^2\) Data showing that physiological and genetic risk factors for mental illness extend across, rather than are bound by, traditional diagnoses,\(^9\) further supports this paradigm, as does the frequent prescription of psychotropic medications for off-label use across diagnostic boundaries.\(^2\) Transdiagnostic studies are also superior to single-diagnosis case–control studies in that they can determine which relationships are shared across various diagnoses and which are unique to a particular disorder.

In this study, we sought to determine whether (1) neurocognition, (2) core clinical dimensions and (3) alcohol and substance use are associated with social and occupational functioning and the magnitude of these associations. The rationale for examining clinical symptoms and functioning alongside neurocognition, sleep changes and substance use is underscored by a recent systematic review highlighting the transdiagnostic relevance of these key domains in youth with mental illness.\(^2\)

In keeping with prior research,\(^7\)\(^1\)\(^2\)\(^5\) it was hypothesised that neurocognition and negative symptoms would make the greatest contribution to level of social and occupational functioning irrespective of the cross-sectional diagnosis applied to cases at the time of inclusion in the cohort. Given the high degree of heterogeneity expected in a transdiagnostic youth sample, we secondarily sought to determine the influence of demographic (eg, age, gender) and clinical factors (eg, diagnosis, medication exposure) on findings.

**METHODS**

Data included in the current study represent the baseline assessments conducted at entry to the cohort study and were collected between April 2008 and May 2016.

**Participants**

Participants were recruited into the Brain and Mind Youth Cohort from youth mental health outpatient services at the Brain and Mind Centre.\(^2\)\(^4\)\(^2\)\(^5\) Referred participants were 12–36 years of age and presented with a major affective, psychotic or developmental/behavioural syndrome. Participants were excluded if they (or their guardians, if aged under 16 years) were unwilling or unable to provide written informed consent, or if they had a pre-existing neurological condition, clinically assessed impaired English language skills and/or intellectual disability.
that precluded completion of study self-ratings. Eligible participants completed a series of observer and self-rated questionnaires.

**Procedure**

Treating clinicians recorded clinical diagnoses, and these were reviewed at consensus meetings by senior, treating psychiatrists (eg, IBH, EMS) and formal diagnostic Epilepsy (n=26), brief psychotic disorder (n=11), substance-induced psychotic disorder (n=14), psychotic disorder not otherwise specified (n=74). §Panic disorder (n=4), social phobia (n=29), obsessive-compulsive disorder (n=11), post-traumatic stress disorder (n=5), generalised anxiety disorder (n=60). †Asperger’s disorder (n=16), attention-deficit/hyperactivity disorder (n=47), conduct disorder (n=7), oppositional defiant disorder (n=4). **Medication data were available in 877 individuals (87.4%), with missing data for the typologies of depression (n=64), bipolar (n=13), psychosis (n=12), anxiety (n=28), developmental (n=9). AUDIT, Alcohol Use Disorders Identification Test; BPRS, Brief Psychiatric Rating Scale; dev/behav, developmental/behavioural; HDRS, Hamilton Depression Rating Scale; SOFAS, Social and Occupational Functioning Assessment Scale; WHO-ASSIST, WHO–Alcohol, Smoking and Substance Involvement Screening Test.

**Table 1** Demographic, clinical and functional characteristics across diagnostic subgroups

|                      | Depression* (n=449) | Bipolar† (n=178) | Psychosis ‡ (n=193) | Anxiety§ (n=109) | Dev/behav¶ (n=74) |
|----------------------|---------------------|------------------|---------------------|------------------|------------------|
|                      | M ± SD              | M ± SD           | M ± SD              | M ± SD           | M ± SD           |
| Age                  | 19.8 ± 4.3          | 21.6 ± 4.8       | 22.2 ± 4.6          | 19.9 ± 4.8       | 17.0 ± 4.6       |
| Education (years)    | 11.6 ± 2.4          | 12.3 ± 2.2       | 12.0 ± 2.4          | 11.5 ± 2.7       | 10.0 ± 2.8       |
| BPRS depression (/7) | 2.4 ± 0.8           | 2.2 ± 0.8        | 2.1 ± 0.9           | 2.3 ± 0.9        | 1.7 ± 0.7        |
| BPRS mania (/7)      | 1.3 ± 0.5           | 1.5 ± 0.6        | 1.4 ± 0.4           | 1.4 ± 0.4        | 1.5 ± 0.7        |
| BPRS positive (/7)   | 1.3 ± 0.4           | 1.4 ± 0.5        | 1.8 ± 0.7           | 1.4 ± 0.5        | 1.3 ± 0.4        |
| BPRS negative (/7)   | 1.5 ± 0.6           | 1.3 ± 0.5        | 1.9 ± 0.8           | 1.5 ± 0.7        | 1.5 ± 0.6        |
| BPRS disorientation (/7) | 1.2 ± 0.5    | 1.1 ± 0.5        | 1.2 ± 0.6           | 1.2 ± 0.5        | 1.2 ± 0.5        |
| HDRS sleep (/6)      | 2.0 ± 1.8           | 1.7 ± 1.7        | 1.3 ± 1.6           | 1.5 ± 1.5        | 1.8 ± 1.7        |
| AUDIT alcohol use (/40) | 6.8 ± 7.4  | 9.0 ± 8.4       | 6.2 ± 8.0           | 4.4 ± 6.4        | 5.1 ± 7.5        |
| WHO-ASSIST tobacco use (/4) | 1.3 ± 1.6  | 1.6 ± 1.7       | 1.7 ± 1.9           | 1.0 ± 1.5        | 1.4 ± 1.7        |
| WHO-ASSIST cannabis use (/4) | 0.7 ± 1.2  | 0.7 ± 1.3       | 0.4 ± 1.0           | 0.4 ± 0.9        | 0.7 ± 1.3        |
| WHO-ASSIST other illicit substance use (/4) | 0.1 ± 0.3 | 0.1 ± 0.3       | 0.1 ± 0.2           | 0.1 ± 0.3        | 0.1 ± 0.3        |
| SOFAS                | 61.8 ± 10.7         | 63.9 ± 11.5      | 55.9 ± 12.1         | 63.2 ± 11.1      | 61.7 ± 9.8       |
| Gender (female)      | 277 ± 61.7          | 129 ± 72.5       | 58 ± 30.1           | 58 ± 53.2        | 20 ± 27.0        |
| Medicated**          | 231 ± 60.0          | 128 ± 77.6       | 140 ± 77.3          | 37 ± 45.7        | 32 ± 49.2        |
| Antidepressants      | 201 ± 52.2          | 63 ± 38.2        | 55 ± 30.4           | 25 ± 30.9        | 12 ± 18.5        |
| Lithium/anticonvulsants | 28 ± 7.3  | 69 ± 41.8        | 24 ± 13.3           | 6 ± 7.4          | 3 ± 4.6          |
| Antipsychotics       | 71 ± 18.4           | 77 ± 46.7        | 125 ± 69.1          | 9 ± 11.1         | 8 ± 12.3         |
| Stimulants           | 14 ± 3.6            | 7 ± 4.2          | 3 ± 1.7             | 4 ± 4.9          | 14 ± 21.5        |

*Major depressive disorder (n=313), dysthymic disorder (n=4), depressive disorder not otherwise specified (n=132). †Bipolar I disorder (n=13), bipolar II disorder (n=25), cyclothymic disorder (n=1), bipolar disorder not otherwise specified (n=139). ‡Schizophrenia (n=53), schizoaffective disorder (n=15), schizoaffective disorder (n=26), brief psychotic disorder (n=11), substance-induced psychotic disorder (n=14), psychotic disorder not otherwise specified (n=74). §Panic disorder (n=4), social phobia (n=29), obsessive-compulsive disorder (n=11), post-traumatic stress disorder (n=5), generalised anxiety disorder (n=60). ¶Asperger’s disorder (n=16), attention-deficit/hyperactivity disorder (n=47), conduct disorder (n=7), oppositional defiant disorder (n=4). **Medication data were available in 877 individuals (87.4%), with missing data for the typologies of depression (n=64), bipolar (n=13), psychosis (n=12), anxiety (n=28), developmental (n=9). AUDIT, Alcohol Use Disorders Identification Test; BPRS, Brief Psychiatric Rating Scale; dev/behav, developmental/behavioural; HDRS, Hamilton Depression Rating Scale; SOFAS, Social and Occupational Functioning Assessment Scale; WHO-ASSIST, WHO–Alcohol, Smoking and Substance Involvement Screening Test.

Board-certified treating clinicians (ie, consultant psychiatrists, clinical psychologists, mental health nurses) provided an evaluation of each participant’s social and occupational functioning. Next, board-certified clinical psychologists, clinical neuropsychologists or trained research psychologists (ie, graduate-level academic psychologists, supervised by RSCL to ensure a sufficient level of inter-rater reliability) conducted structured clinical interviews, neuropsychological testing, as well as an additional assessment of social and occupational functioning to improve the reliability of this single, clinician-rated score (approximately 80% were conducted within a month of the treating clinician assessment).
Alcohol and substance use

► Social and occupational functioning was indexed using the Social and Occupational Assessment Functioning Scale (SOFAS). Scores were averaged across the treating clinician and researcher assessments (intraclass correlation coefficient=0.70), as previously done. This composite score was derived to obtain a more reliable estimate of real-world functioning and, secondarily, to conserve free parameters and increase stability of parameter estimates. A higher score denotes better functioning.

► Neurocognition was assessed using a broad neuropsychological battery with demographic normative-adjustments (previously described) and was chosen on the basis of sound psychometric properties and relevance to the disorders under study. Predicted IQ was estimated using the Wechsler Test of Adult Reading or Wide Range Achievement Test—fourth edition (for participants younger than 16 years). Psychomotor speed and mental flexibility were measured using Trail Making Test—Part A (TMT-A) and Part B (TMT-B). Verbal learning and memory were indexed using the five-trial total and delayed recall scores from the Rey Auditory Verbal Learning Test. Verbal fluency was composed of the letter (F-A-S) and category (animals) fluency subsets of the Controlled Oral Word Association Test. A higher score indicates better functioning.

► Core clinical symptom dimensions were measured across two validated scales. Symptoms were rated on the expanded Brief Psychiatric Rating Scale (BPRS) using empirically derived symptom subscores (depression and anxiety, mania, positive symptoms, negative symptoms and disorientation). The BPRS does not capture sleep profiles as a separate dimension so, as in previous studies, disturbed sleep was indexed using the sum of the three sleep items from the Hamilton Depression Rating Scale. A higher score denotes greater severity of symptoms.

► Alcohol and substance use were measured across two validated scales. Alcohol use was indexed using the Alcohol Use Disorders Identification Test total score. Substance use for tobacco, cannabis and other illicit substances was measured using the ‘current frequency’ subscale (past 3 months) from the WHO–Alcohol, Smoking and Substance Involvement Screening Test questionnaire. A higher score indicates greater alcohol or substance use.

Patient and public involvement

Participants were not involved in the development of research question(s), design and outcome measures, nor was the study informed by their priorities, experience and preferences. We did not formally assess the burden of time required to participate in the research.

Data analysis

Statistical analyses were conducted using IBM SPSS V.20.0 and AMOS V.20.0. Maximum likelihood estimation (MLE) was employed for all structural equation modelling (SEM) analyses. MLE was chosen as it is the most robust approach in the potential event of statistical assumption violations and performs best in heterogeneous samples. Missing data were also handled by MLE which does not involve data imputation, but uses all available data to compute maximum likelihood estimates. Diagnostic and demographic data were available for all participants. In total, 9.1% of data were missing for functioning, 18.8% for neurocognition, 12.7% for clinical symptoms and disturbed sleep, and 17.7% were missing for alcohol and substance use. Each analysis had >80% of cases with complete data. Additional analyses revealed that data were not missing at random, and missing data were more likely to occur in younger participants (Welch’s F(1102.54)=4.85, p<0.05) and in those with an anxiety disorder (χ²(4)=26.09, p < 0.001), although the effect sizes were small (Cohen’s d=0.26 and Cramer’s V=0.16, respectively).

Normality was assessed through inspection of Q-Q plots, given inferential measures of non-normality (eg, Shapiro-Wilk statistic) are overly sensitive in large datasets and almost always return a significant finding. All endogenous variables (eg, SOFAS) met normality assumptions on visual inspection. Based on visual inspection of the frequency histograms and assessment of the Q-Q plot, the predictor/exogenous variables that departed from normality were positive symptoms, negative symptoms, mania, disorientation, TMT-A and TMT-B which were all observed to have a slight positive skew (no others were skewed). Prior studies have found that MLE methods are robust in cases where variables depart from normality where n>600, as in the present case. However, other approaches to non-normal data, such as asymptotically distribution free SEM, require no missing data and would unsatisfactorily affect the generalisability of findings as well as statistical power in the current analyses. As such, we used the MLE approach.

We first used SEM to evaluate the best-fitting measurement model for the following predictors: (1) neurocognition, (2) clinical symptoms and disturbed sleep, and (3) alcohol and substance use. Then, we used SEM to test the structural model (ie, the relationship between predictors and social and occupational functioning) at both the single-predictor and the overall levels in order to explore potential predictors and delineate unique contributions. This was done in a two-step process—first, by testing individual predictors and then by testing the combined predictors—to quantify the amount of overlapping and unique explanatory power. All analyses used a model-trimming approach through an iterative process in which non-significant paths with the smallest contribution were sequentially eliminated from a saturated model (where all variables were allowed to freely co-vary), until a best fitting model was derived to best explain the relationships between predictors and functional impairment. Finally, modification indices generated by AMOS were used to optimise model fit (ie, to inform which participants' responses were not influenced by their priorities, experience, and preferences. We did not formally assess the burden of time required to participate in the research.
paths and parameters should be added or removed to increase model adequacy), although these were only used when deemed theoretically meaningful. Residuals were allowed to correlate if theoretically justified (eg, common measurement variance between neuropsychological subtests).

Model fit was determined using: (1) the absolute fit $\chi^2$ statistic and (2) the relative fit indices: Bentler Comparative Fit Index (CFI,\textsuperscript{39} Bentler-Bonnett Non-normed Fit Index (NFI\textsuperscript{40} and Root Mean Square Error of Approximation (RMSEA\textsuperscript{41} with 90% CI. An excellent-fitting model is typically indicated by a non-significant $\chi^2$ test (indicating a non-significant difference between the covariance matrix of the data and the model), a CFI and NFI of greater than 0.90 (indicating that the current model was superior to a null model where all paths are constrained to zero), and a RMSEA of less than 0.05 with an upper CI bound of less than 0.08 (indicating that the error of approximation and a RMSEA of less than 0.05 with an upper CI bound of less than 0.08 (indicating that the error of approximation of the model compared with the data was acceptable). In small samples (ie, less than 200), the $\chi^2$ statistic has been shown to be an adequate index of absolute model fit.\textsuperscript{36} However, as sample size increases, the $\chi^2$ statistic (relative to a constant df) disproportionally increases, and is nearly always significant and inappropriately rejects the model irrespective of specified parameters.\textsuperscript{42 43} An alternative solution is to compute a relative $\chi^2$/df ratio, with a value between 2 and 5 considered excellent to adequate fit.\textsuperscript{42 44–46} although primary emphasis will be placed on the relative fit indices as is the established convention.\textsuperscript{13}

Moderator analyses were conducted allowing a model to be tested in separate subgroups, comparing the parameter estimates to determine how predictors of social and occupational functioning in the final model are moderated by demographic and clinical factors (these were dichotomous to maintain statistical power within subgroups for this categorical procedure). For instance, the median-split on age was performed to determine whether the model held for both younger and older individuals while maintaining statistical power. We sought to specifically test whether predictors in affective-spectrum disorders (anxiety, depressive and bipolar disorders) were similarly associated with functional impairment compared with psychotic, developmental or behavioural conditions. We chose to include primary affective disorders (ie, major depression, bipolar disorder or an anxiety disorder) as a moderator since these disorders have been shown to carry less neurocognitive burden in recent-onset mood disorders\textsuperscript{47 48} and, as such, could potentially influence the role of neurocognition and the magnitude of effects in the statistical models.

RESULTS
Sample characteristics
In total, 1003 patients were recruited. As shown in tables 1 and 2, cross-sectional diagnoses were composed of depressive (n=449), bipolar (n=178), psychotic (n=193), anxiety (n=109), and developmental or behavioural disorders (n=74). The mean age was 20.4 years (SD=4.7), with 54.0% being female (n=542). Mean educational attainment was 11.7 years (SD=2.5), with an average predicted IQ of 101.9 (SD=10.8). The mean SOFAS score was 61.2 (SD=11.4), indicating moderate levels of impairment. Of the participants with medication data available (87.4%, n=877), 64.8% were prescribed psychotropic medications (n=568). Of these 568 cases, 40.6% were prescribed an antidepressant (n=356), 14.8% were prescribed lithium or an anticonvulsant (n=130), 33.1% were prescribed any antipsychotic (n=290) and 4.8% were prescribed a stimulant (n=42).

Single-predictor models
A. Neurocognition. Inspecting the screen plot, exploratory factor analyses identified two potential latent structures. The one-factor model was a very good fit for the data ($\chi^2=57.3$, df=17, $p<0.001$, CFI=0.980, NFI=0.972, RMSEA=0.055, 90% CI 0.040 to 0.071), and was a better fit than a two-factor model, whereby trails A, trails B and IQ loaded on one latent factor, and IQ, Rey total, Rey delay, FAS and animals loaded on a second

| Table 2 | Neuropsychological functioning across diagnostic subgroups |
|---------|---------------------------------------------------------|
| **Depression** (n=449) | **Bipolar** (n=178) | **Psychosis** (n=193) | **Anxiety** (n=109) | **Dev/behav** (n=74) |
| M     | SD    | M     | SD    | M     | SD    | M     | SD    | M     | SD    |
| IQ$^*$ | 103.25| 10.49 | 102.82| 9.08  | 99.99 | 10.47 | 102.66| 9.59  | 95.30 | 14.83 |
| Trails A† | –0.01 | 1.32 | 0.16  | 1.06  | –0.31 | 1.01  | 0.12  | 0.96  | –0.03 | 1.07  |
| Trails B† | –0.44 | 1.53 | –0.47 | 1.92  | –1.22 | 2.23  | –0.50 | 1.59  | –0.85 | 2.16  |
| Rey total† | –0.06 | 1.27 | –0.18 | 1.20  | –1.12 | 1.46  | 0.09  | 1.93  | –0.48 | 1.42  |
| Rey delay† | 0.03  | 1.34 | –0.27 | 1.38  | –1.07 | 1.51  | 0.18  | 2.29  | –0.29 | 1.22  |
| FAS† | –0.31 | 1.15 | –0.04 | 1.07  | –0.56 | 1.00  | –0.37 | 1.15  | –0.86 | 1.09  |
| Animals† | 0.23  | 1.20 | 0.45  | 1.25  | –0.28 | 1.08  | 0.26  | 1.17  | –0.03 | 0.94  |

$^*$Age-adjusted; normative M=100; SD=15.
†Demographically adjusted; normative M=0.00; SD=1.00.
Dev/behav, developmental/behavioural.
latent factor ($\chi^2=76.499$, df=11, $p<0.001$, CFI=0.967, NFI=0.962, RMSEA=0.077, 90% CI 0.061 to 0.094). Factor loadings on the one-factor model were all significant and ranged from 0.51 to 0.69 (see figure 1A). Neurocognition was a significant contributor to functional level ($\beta=0.39$, $p<0.001$), explaining 15% of the variance.

B. Core clinical symptom dimensions. Only three clinical dimensions (depression and anxiety ($\beta=-0.18$, $p<0.001$), positive symptoms ($\beta=-0.17$, $p<0.001$) and negative symptoms ($\beta=-0.26$, $p<0.001$)) were associated with functioning, whereas mania and disorientation were not significantly associated ($p$ values $>0.05$). The model demonstrated excellent fit ($\chi^2=8.6$, df=8, $p=0.379$, CFI=0.999, NFI=0.986, RMSEA=0.009, 90% CI 0.000 to 0.043), with the three dimensions explaining a total of 18% of the variance in functioning (see figure 1B).

C. Alcohol and substance use. Exploratory factor analyses determined that alcohol use did not load with the other substance use variables. Only a two-factor latent model was possible given the number of observed variables and statistical constraints. The two-factor model emerged as an excellent fit for the data ($\chi^2=7.4$, df=4, $p=0.116$, CFI=0.995, NFI=0.990, RMSEA=0.033, 90% CI 0.000 to 0.069), whereby tobacco, cannabis and other illicit substance use loaded on a single ‘substance use’ latent variable as distinct from alcohol use (figure 1C). Only substance use was predictive of functioning ($\beta=-0.10$, $p<0.05$), explaining 1% of the variance.

Final model
In the overall model, all the factors identified in the single predictor models remained significant, except for substance use (figure 2). Neurocognition showed the strongest unique contribution to social and occupational functioning ($\beta=0.36$, $p<0.001$); depressive symptoms were next ($\beta=-0.24$, $p<0.001$), followed by negative symptoms ($\beta=-0.15$, $p<0.001$) and finally positive symptoms ($\beta=-0.10$, $p<0.001$). Together, these four clinical features independently accounted for 31% of the variance in functioning, with the final model being a very good fit for the data ($\chi^2=279.8$, df=119, $p<0.000$, CFI=0.956, NFI=0.926, RMSEA=0.037, 90% CI 0.031 to 0.042). Mania, disorientation and alcohol and substance use all significantly correlated with these four significant features ($p$ values $<0.05$).

Moderator analyses
► Age. As shown in table 3, positive symptoms were no longer a significant contributor to functioning in the
12-year-old to 20-year-old group ($\beta=-0.06$, $p=0.178$). The model with older individuals explained 18% more variance in functional impairment than the model with younger individuals. This was driven in large part by a difference in predictive strength of neurocognition, whereby it was more predictive in older ($\beta=0.44$, $p<0.001$) than younger individuals ($\beta=0.27$, $p<0.001$).

► Gender. Positive symptoms were non-significant in the male subgroup ($\beta=-0.07$, $p=0.145$), whereas all other clinical features remained significant ($p$ values <0.001). In females, negative symptoms became non-significant ($\beta=-0.07$, $p=0.105$), while the other contributors remained significant. The final model was comparable across genders in terms of the total variance explained.

► Primary affective disorder diagnosis. Neurocognition, depression and anxiety, and negative symptoms remained significant contributors to functional level irrespective of affective disorder diagnosis (see table 4). By contrast, positive symptoms no longer remained significant in both the affective disorder ($\beta=-0.63$, $p=0.097$) and psychosis, developmental or behavioural disorders ($\beta=-0.102$, $p=0.123$) subgroups. An additional 14% of the variance in functioning was explained in individuals with a psychotic, developmental or behavioural disorder, primarily owing to the greater predictive strength of neurocognition ($0.30$ vs $0.43$, $p<0.001$).

► Medication usage. All factors associated with functional impairment remained significant in participants who were unmedicated. By contrast, positive symptoms no longer remained significant in medicated individuals ($\beta=-0.06$, $p=0.117$).

Sensitivity analysis

Restricting the full sample to individuals aged 15–25 years (n=794) yielded a very good fitting model as well ($\chi^2=240.1$, df=119, $p<0.000$, CFI=0.959, NFI=0.924, RMSEA=0.036, 90% CI 0.029 to 0.042). The explained variance remained the same (31% explained). Importantly, all predictors remained significant with the same effect sizes, with the exception of depression and anxiety, which became slightly more predictive ($-0.25 \rightarrow -0.26$), and neurocognition, which became slightly less predictive ($0.36 \rightarrow 0.35$).

Table 3 Analyses of demographic factors (age, gender) as moderators of the relationships between predictors and functional outcome in the final model

| Age* | Gender† |
|------|---------|
| 12–20 Years (n=539) | 21–36 Years (n=464) | Male (n=461) | Female (n=542) |
| Neurocognition | $\beta$ | P values | $\beta$ | P values | $\beta$ | P values | $\beta$ | P values |
| Depression and anxiety | $-0.28$ | 0.000 | $-0.22$ | 0.000 | $-0.23$ | 0.000 | $-0.30$ | 0.000 |
| Positive symptoms | $-0.06$ | 0.178 | $-0.14$ | 0.002 | $-0.07$ | 0.145 | $-0.12$ | 0.004 |
| Negative symptoms | $-0.13$ | 0.005 | $-0.18$ | 0.000 | $-0.19$ | 0.000 | $-0.07$ | 0.105 |

*K20 Years (subgroup Model, $R^2=0.24$); 21–36 Years (Subgroup Model, $R^2=0.40$).
†Male (subgroup model, $R^2=0.32$); female (subgroup model, $R^2=0.29$).
Table 4 Analyses of clinical factors (primary affective disorder, medication usage) moderating the relationship between predictors and functional outcome in the final model

| Primary affective disorder* | Medication usage† |
|---------------------------|-------------------|
|                          | Yes (n=736)       | No (n=267)       | Nil (n=309)       | Medicated (n=568) |
|                          | β  P values       | β  P values      | β  P values       | β  P values       |
| Neurocognition            | 0.30 0.000        | 0.43 0.000       | 0.38 0.000        | 0.38 0.000        |
| Depression and anxiety    | −0.29 0.000       | −0.24 0.000      | −0.15 0.007       | −0.24 0.000       |
| Positive symptoms         | −0.06 0.097       | −0.10 0.123      | −0.22 0.000       | −0.06 0.117       |
| Negative symptoms         | −0.12 0.003       | −0.16 0.009      | −0.19 0.000       | −0.13 0.002       |

*Yes (subgroup model, $R^2=0.24$); no (subgroup model, $R^2=0.38$).
†Nil (subgroup model, $R^2=0.38$); medicated (subgroup model, $R^2=0.29$).

DISCUSSION

In a large, clinical, transdiagnostic cohort of youth with mental disorders, impaired neurocognition was the clinical feature most significantly associated with functional impairment. The role of neurocognition was attenuated but still significant in those with an affective disorder diagnosis and in the younger age group. The findings are relevant as they demonstrate that while neurocognitive impairment may undermine functioning in those with psychotic disorders, or in chronic or recurrent mental disorders, they are not specific to such cases. That is, neurocognitive dysfunction has traditionally been argued as a core, underlying feature of social and occupational impairments in chronic schizophrenia. However, our current findings support the burgeoning position that the role of neurocognitive deficits cuts across diagnosis and clinical stage. Nevertheless, it appears that neurocognitive disturbances are more pronounced in those with psychotic, developmental or behavioural disorders, converging with evidence of more pronounced cognitive deficits in children who will go on to develop psychosis compared with those who develop depression or bipolar disorder.3,49–51 Mechanistically, whether neurocognitive dysfunction drives functional impairment as a few past studies have found,7 28 and is conversely a consequence of poor functioning remains to be clarified.

Depressive, anxiety and negative symptom dimensions also contributed significantly to level of social and occupational functioning, supporting previous disorder-specific research.11 13 28 Importantly, the contributions of these factors to level of functioning were largely independent of one another, and do not appear to be moderated by other clinical or demographic factors. By comparison, the role of positive symptoms diminished considerably in the final model. This finding differs from other research in psychotic and bipolar disorders, and may reflect the lower prevalence of positive symptoms in our cohort in contrast to previous studies.11 13 28 However, it was notable that positive symptoms in older, unmedicated females remained significantly associated with functioning. As with neurocognition however, the directionality of findings remains unclear, with some evidence suggesting that it may be bidirectional in the case of negative symptoms.7 28 52

Intriguingly, neither alcohol and substance use, nor sleep disturbances, were directly associated with functional impairment, although these factors remained significantly associated with neurocognition and clinical symptoms. Therefore, their role in social and occupational functioning does not appear to be direct, but may operate indirectly (eg, substance use may impair cognition which in turn may impair functioning). The indirect effects of alcohol and substance use, as well as sleep and circadian disruptions, warrant more detailed examination and causal analysis in longitudinal datasets. Moreover, the lack of a direct association between alcohol and substance use and functioning may be related to the domains of neurocognitive functions currently tested. That is, the impact of substance use on functioning may be greatest in other neurocognitive functions that are more directly linked to driving and maintaining alcohol and substance use behaviours, such as those subserved by the fear, reward and self-control circuitries not covered in the current neuropsychological battery (eg, reward-related cue learning, habit formation, response inhibition).

The current findings have important implications for the transdiagnostic, dimensional approach to psychiatry. Research examining the underlying mechanisms of functional impairment in single-diagnosis or dual-diagnosis cohorts have been unable to capture the unique contributions of a comprehensive range of neurocognitive, symptom, sleep and circadian factors, as well as other psychoactive exposures (ie, substance use, prescribed medications).22 In particular, neuropsychological studies in psychosis have not routinely and concurrently assessed depression and anxiety symptoms, hypomania and full-threshold mania, substance misuse and sleep disturbance. That is not to say that categorical, nosological approaches have had little to contribute to the field. Indeed, the key argument underpinning a DSM approach is to allow for comparability across studies and so diagnostic determinations are often necessary. However, in youth, diagnoses tend to be unstable1 and, as such, not as useful. One plausible way forward for dimensional psychiatry is to ensure that the samples used in transdiagnostic studies are characterised as clearly and as comprehensively as possible,16 53 54 as was attempted in the present investigation.
In terms of limitations, the current analyses were cross-sectional, and future research investigating moderator and mediator analyses would benefit from cross-lagged, longitudinal path modelling to disentangle causality.55 Second, the measures used to index clinical symptoms, sleep disturbance, and alcohol and substance use were not as comprehensive as is typical in the sleep and addiction literatures, and some were not originally designed for use in certain clinical disorders which may have reduced sensitivity to detect symptoms (e.g., mania). More detailed examination of these dimensions in the future will help more definitively determine whether the impact of neurocognition on functioning is as large as currently identified. Future studies would also benefit from using real-time, ecological momentary assessment technologies (e.g., substance use monitoring using smartphones, actigraphy monitoring of physical activity and sleep quality). Third, medication data were not available for the full sample (12.6% were missing) and, as such, the moderating role of medication status requires further corroboration (as with the role of medication type). Fourth, clinical diagnoses assigned to cases in the current study were by treating psychiatrists and future studies should consider more structured approaches (e.g., structured clinical interview for DSM), including consideration of the influence of other comorbid diagnoses (e.g., personality disorders). Further, the age range included in the current study meant that individuals on the opposite ends of the age spectrum were at different stages of their cognitive and emotional development (e.g., executive functioning, emotional regulation), although our sensitivity analyses support the argument that our findings hold irrespective of age. Further, a phenotype-approach, as attempted in the current study, would necessarily require converging genetic and neuroimaging evidence to ensure that the neurocognitive and symptom dimensions identified as predictive of functioning are linked to specific neural circuitries (e.g., corticobasal ganglia systems56) and genotype which would ultimately facilitate the development of next-generation and neuroscience-informed pharmacotherapies. Finally, it remains to be seen whether the current findings hold in future studies with less missing data, as well as in studies using measures or approaches that can circumvent potential biases stemming from non-normally distributed data.

This was the first study to examine a broad range of illness-related factors and associations with functional impairment in a well-powered and broadly transdiagnostic, clinical cohort of more than one thousand young people with mental illness. A significant contribution of the present findings to the established literature was evidence showing that neurocognition is a strong and reliable, unique predictor of social and occupational functioning irrespective of diagnosis—in a cohort predominantly composed of affective disorders which has not been previously demonstrated before at this scale. As such, the functional importance of neurocognitive functions clearly extends beyond the psychosis and developmental disorders spectrums and appears to become more pronounced with increasing age. Future studies should attempt to replicate these findings, as well as to clarify the directions of cause and effect.

Author affiliations
1Brain and Mind Centre, University of Sydney, Sydney, New South Wales, Australia
2Brain and Mental Health Research Hub, Monash University, Melbourne, Victoria, Australia
3Charles Perkins Centre, University of Sydney, Sydney, New South Wales, Australia
4The Monash Alfred Psychiatry Research Centre, Monash University, Melbourne, Victoria, Australia
5Marie Bashir Institute, University of Sydney, Sydney, New South Wales, Australia
6Academic Psychiatry, Institute of Neuroscience, Newcastle University, Newcastle, UK
7Centre for Affective Disorders, Institute of Psychiatry, Psychology and Neuroscience, London, UK

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