The Better Half of Selling Separately*

Sergiu Hart† Philip J. Reny‡

November 24, 2021

Abstract

Separate selling of two independent goods is shown to yield at least 62% of the optimal revenue, and at least 73% when the goods satisfy the Myerson regularity condition. This improves the 50% result of Hart and Nisan (2017, originally circulated in 2012).

Contents

1 Introduction 2
1.1 Some Related Work ........................................ 5
1.2 Organization of the Paper .................................... 2

---

*Previous versions: December 2017 (https://arxiv.org/abs/1712.08973), November 2016. We thank Noam Nisan for useful discussions, and the editors and referees for their helpful suggestions and comments.

†Department of Economics, Institute of Mathematics, and Federmann Center for the Study of Rationality, The Hebrew University of Jerusalem. E-mail: hart@huji.ac.il Web site: http://www.ma.huji.ac.il/hart

‡Department of Economics, University of Chicago. Research partially supported by National Science Foundation grant SES-1724747. E-mail: preny@uchicago.edu Web site: http://economics.uchicago.edu/directory/philip-j-reny
1 Introduction

One of the most celebrated aspects of Myerson’s (1981) optimal auction result is that it provides an economic explanation for the ubiquitous use of the four standard auction forms. Strictly speaking, however, Myerson’s results apply only to cases in which a seller is selling a single good. Because many sellers sell multiple goods, extending Myerson’s analysis to the multi-good case has long been considered a critical next step. But the multi-good monopoly problem has resisted a complete solution for over 35 years and by now it is well understood to be an extremely difficult problem. Worse still, it is known that the optimal solution must typically be quite complex and very
often requires buyers to purchase randomized contracts. And therein lies the
difficulty, because we do not often, if ever, observe complex or randomized
selling mechanisms in practice. This raises the obvious question, Why not?

One reason that we may not observe the kinds of complex mechanisms
that full optimality dictates is that relatively simple mechanisms may suffice
for generating much of the revenue that could ever be generated. Thus, a
complementary approach to the research program initiated in Myerson (1981)
is to search for relatively simple mechanisms that yield a significant fraction
of the revenue that is generated by a fully optimal mechanism. The present
paper represents a modest contribution to this program.

We consider a single seller who has one unit each of two indivisible goods.
The two goods need not be identical. The seller, whose value for the two
goods is zero, can offer to sell the goods to a single buyer. The buyer’s two
values, one value for each of the two goods, are unknown to the seller but
are known to be independently distributed (so we say that the two goods
are independent). The buyer is risk neutral and has preferences that are
additive in the values and (negatively so) in the price paid. Even in this
most basic case, there is no known characterization of the optimal selling
mechanism, though it is known that the optimal mechanism can display
unusual properties. We ask, What fraction of the optimal revenue can the
seller guarantee by selling each of the two goods separately, i.e., by posting
a Myerson-optimal price for each of the goods?

In the context of a general analysis with any finite number of indepen-
dent goods, Hart and Nisan (2017, originally circulated in 2012) show in
particular that, by selling two independent goods separately, the seller can
guarantee at least 50% of the optimal revenue but cannot guarantee more

---

1For examples of this approach, see Hart and Nisan (2017) and the references therein.
2For example, the optimal mechanism can be non-monotonic (i.e., increasing the buyer’s
valuations may well decrease the seller’s optimal revenue), and it can require the buyer
to accept randomized contracts. See Hart and Reny (2015); Section 3 there contains
references to previous examples that require such randomizations.
than 78%. A nice feature of the 50% revenue guarantee is that its proof is relatively simple. In part, this simplicity arises from the rather generous bounds that are established at various steps. While it seems clear that the bounds employed in the Hart–Nisan proof are “much” too generous, tightening them as we do here requires a surprising amount of additional effort. Hart and Nisan also show that if, in addition, the buyer’s two independent values are identically distributed, then the revenue guarantee is at least 73%, which is tantalizingly close to the 78% upper limit.

Our main result significantly improves upon the Hart–Nisan 50% guarantee, and shows that their 73% guarantee with identically distributed values can also be obtained when the buyer’s value distributions satisfy Myerson-regularity. None of our results require the two values to be identically distributed.

**Main Result.** For any two independent goods, selling each good separately at its optimal one-good price guarantees at least \( \sqrt{e}/(\sqrt{e} + 1) \approx 62\% \) of the optimal revenue. Furthermore, if the buyer’s two value distributions each satisfy Myerson-regularity, then the guaranteed fraction of optimal revenue increases to \( e/(e + 1) \approx 73\% \).

This is stated below as Theorems 7 and 9.

To summarize the known bounds on the guaranteed fraction of optimal revenue (GFOR) from selling separately two goods: when the goods are independent, the GFOR is at least 62%; when they are independent and either Myerson-regular or identically distributed, the GFOR is at least 73%; in all these three cases, the GFOR is at most 78%; and, when the goods are not necessarily independent, the GFOR drops all the way down to zero (Hart and Nisan 2013).³

³Hart and Nisan (2017) establish the 78% upper bound with an explicit example in which it is optimal to sell two independent and identically distributed goods as a bundle (these goods satisfy the Myerson-regularity condition).

⁴For more than two goods a similar result is due to Briest, Chawla, Kleinberg, and
1.1 Some Related Work

There is by now a vast literature in game theory, economics, and computer science that deals with the (optimal) selling of multiple goods. While that literature is too large to survey here, the reader may wish to consult the literature section in, say, Hart and Nisan (2017) for an overview. We will mention here the work of Babaioff, Immorlica, Lucier, and Weinberg (2014) that shows that the better option between selling the goods separately and selling them as the bundle of all goods yields a GFOR that is bounded away from zero for any number of goods. Recently, in the case of two independent goods, Babaioff, Nisan, and Rubinstein (2018) have shown that separate selling yields at least 78% of the optimal deterministic revenue, and that this bound is tight. In the related setup of a unit-demand buyer (who desires to buy only one good, rather than having an additive value over bundles of goods), Chawla, Malec, and Sivan (2010, Theorem 5) show a GFOR of 1/4 for the separate selling of any number of independent goods. Finally, Daskalakis, Deckelbaum, and Tzamos (2017) provide a useful duality characterization of the revenue-optimizing mechanism for multiple goods.

1.2 Organization of the Paper

The paper is organized as follows. Section 2 presents the model, defines the appropriate concepts, and provides some preliminary results. Section 3 gives an outline of the proof. The proof itself consists of a general decomposition result (Proposition 4 in Section 4) and an estimate of the crucial term there (Proposition 6 in Section 4), which, when combined, give the first part of the Main Theorem, namely, the general 62% bound (Theorem 7 in Section 6). Section 7 proves the second part of the Main Theorem, namely, the 73% bound for regular distributions (Theorem 9), together with some additional results. Appendix A provides a general result on the continuity of the revenue

Weinberg (2015, originally circulated in 2010).
with respect to valuations (which is of independent interest), and Appendix \[ \text{Appendix} \] gives a simple illustration of the use of nonsymmetric diagonals.

# 2 Preliminaries

## 2.1 The Model

The basic model is standard, and the notation follows Hart and Reny (2015) and Hart and Nisan (2017), which the reader may consult for further details and references.

One seller (or “monopolist”) is selling a number $k \geq 1$ of goods (or “items,” “objects,” etc.) to one buyer. The goods have no value or cost to the seller. Let $x_1, x_2, \ldots, x_k \geq 0$ be the buyer’s values for the goods. The value for getting a set of goods is additive: getting the subset $I \subseteq \{1, 2, \ldots, k\}$ of goods is worth $\sum_{i \in I} x_i$ to the buyer (and so, in particular, the buyer’s demand is not restricted to one good only). The valuation of the goods is given by a random variable $X = (X_1, X_2, \ldots, X_k)$ that takes values in $\mathbb{R}_+^k$ (we thus assume that valuations are always nonnegative); we will refer to $X$ as a $k$-good random valuation. The realization $x = (x_1, x_2, \ldots, x_k) \in \mathbb{R}_+^k$ of $X$ is known to the buyer, but not to the seller, who knows only the distribution $F$ of $X$ (which may be viewed as the seller’s belief); we refer to a buyer with valuation $x$ also as a buyer of type $x$. The buyer and the seller are assumed to be risk neutral and to have quasi-linear utilities.

The objective is to maximize the seller’s (expected) revenue.

By the Revelation Principle (Myerson 1981), it is without loss of generality to restrict attention to “direct mechanisms” that are “incentive compatible.” A direct mechanism $\mu$ consists of a pair of functions $\delta(q, s)$, where

---

\[ ^5 \text{All functions in this paper are assumed to be Borel measurable (cf. Hart and Reny 2015, footnotes 10 and 48).} \]
\( q = (q_1, q_2, \ldots, q_k) : \mathbb{R}_+^k \to [0, 1]^k \) and \( s : \mathbb{R}_+^k \to \mathbb{R} \). If the buyer reports a valuation vector \( x \in \mathbb{R}_+^k \), then \( q_i(x) \in [0, 1] \) is the probability that the buyer receives good \( i \) (for \( i = 1, 2, \ldots, k \)), and \( s(x) \) is the payment that the seller receives from the buyer. When the buyer reports his value \( x \) truthfully, his payoff is \( b(x) = \sum_{i=1}^{k} q_i(x) x_i - s(x) = q(x) \cdot x - s(x) \), and the seller’s payoff is \( s(x) \).

The mechanism \( \mu = (q, s) \) satisfies individual rationality (IR) if \( b(x) \geq 0 \) for every \( x \in \mathbb{R}_+^k \); it satisfies incentive compatibility (IC) if \( b(x) \geq q(\tilde{x}) \cdot x - s(\tilde{x}) \) for every alternative report \( \tilde{x} \in \mathbb{R}_+^k \) of the buyer when his value is \( x \), for every \( x \in \mathbb{R}_+^k \); and it satisfies no positive transfer (NPT) if \( s(x) \geq 0 \) for every \( x \in \mathbb{R}_+^k \) (which, together with IR, implies that \( s(0) = b(0) = 0 \)).

The (expected) revenue of a mechanism \( \mu = (q, s) \) from a buyer with random valuation \( X \), which we denote by \( R(\mu; X) \), is the expectation of the payment received by the seller; i.e., \( R(\mu; X) = \mathbb{E}[s(X)] \). We now define:

- **Rev(\( X \)),** the optimal revenue, is the maximal revenue that can be obtained: \( \text{Rev}(X) := \sup_{\mu} R(\mu; X) \), where the supremum is taken over all mechanisms \( \mu \) that satisfy IR and IC.

When there is only one good, i.e., when \( k = 1 \), Myerson’s (1981) result is that

\[
\text{Rev}(X) = \sup_{p \geq 0} \int X \geq p \mathbb{P}[X \geq p] = \sup_{p \geq 0} \int X > p \mathbb{P}[X > p] = \sup_{p \geq 0} p \cdot (1 - F(p)), \tag{1}
\]

where \( F \) is the cumulative distribution function of \( X \). Optimal mechanisms correspond to the seller “posting” a price \( p \) and the buyer buying the good.

---

6 Without loss of generality any mechanism can always be extended to the whole space \( \mathbb{R}_+^k \); see Hart and Reny (2015).
7 When the goods are infinitely divisible and the valuations are linear in quantities, \( q_i \) may be alternatively viewed as the *quantity* of good \( i \) that the buyer gets.
8 The scalar product of two \( n \)-dimensional vectors \( y = (y_1, \ldots, y_n) \) and \( z = (z_1, \ldots, z_n) \) is \( y \cdot z = \sum_{i=1}^{n} y_i z_i \).
9 Individual rationality recognizes that, regardless of his valuation, the buyer can obtain an expected payoff of zero by not participating in the mechanism.
for the price \( p \) whenever his value is at least \( p \); in other words, the seller makes the buyer a “take-it-or-leave-it” offer to buy the good at price \( p \).

Besides the maximal revenue \( \text{Rev}(X) \), we consider what can be obtained from the simple class of mechanisms that sell each good separately.

- \( \text{SRev}(X) \), the \emph{separate revenue}, is the maximal revenue that can be obtained by selling each good separately. Thus

\[
\text{SRev}(X) = \text{Rev}(X_1) + \text{Rev}(X_2) + ... + \text{Rev}(X_k).
\]

The separate revenue is obtained by solving \( k \) one-dimensional problems (using (I)), one for each good.

We now state the basic properties from Hart and Nisan (2017, Propositions 5 and 6) needed for our proof.

**Proposition 1** (i) Let \( \mu = (q,s) \) be a mechanism for \( k \) goods with buyer payoff function \( b \). Then \( \mu = (q,s) \) satisfies IC if and only if \( b \) is a convex function and for all \( x \) the vector \( q(x) \) is a subgradient of \( b \) at \( x \) (i.e., \( b(\tilde{x}) - b(x) \geq q(x) \cdot (\tilde{x} - x) \) for all \( \tilde{x} \)).

(ii) \( \text{Rev}(X) = \sup_{\mu} R(\mu;X) \) with the supremum taken over all IC, IR, and NPT mechanisms \( \mu \).

### 2.2 Distributions

As we show formally in Appendix A.1 for the results of the present paper we can limit ourselves without loss of generality to valuations that admit a density function (this follows from general continuity properties of the revenue, which we prove in Appendix A and are of independent interest).

In what follows we thus assume that every nonnegative random variable \( X \) has an absolutely continuous cumulative distribution function, \( F(t) =\)
\( \mathbb{P}[X \leq t] = \mathbb{P}[X < t] \), with an associated density function \( f(t) \). We denote by \( G \) the tail probability, i.e.,

\[
G(t) := 1 - F(t) = \int_t^\infty f(u)du = \mathbb{P}[X \geq t],
\]
and by \( H \) the cumulative tail probability, i.e.,

\[
H(t) := \int_0^t G(u)du = \mathbb{E}[\min\{X, t\}]
\]
(2)

(2)

The equality holds since \( \mathbb{E}[\min\{X, t\}] = \int_0^\infty \mathbb{P}[\min\{X, t\} \geq u]du = \int_0^t \mathbb{P}[X \geq u]du = \int_0^t G(u)du \).

Let \( r :\text{REV}(X) > 0 \) be the optimal revenue from \( X \); then (1) implies \( G(t) \leq r/t \), which together with \( G(t) \leq 1 \) gives

\[ G(t) \leq \min\left\{ \frac{r}{t}, 1 \right\}. \]

Therefore

\[
H(t) \leq \int_0^r 1du + \int_r^t \frac{r}{u}du = r + r \log \frac{t}{r},
\]
(3)

(3)

for every \( t \geq r \) (and \( H(t) \leq t \) for \( t \leq r \)).

### 2.3 Change of Units

We start with a trivial, but useful, change of units. For every \( 0 < \lambda_1, ..., \lambda_k \leq 1 \), let \( \mathcal{M}_{\lambda_1, ..., \lambda_k} \) denote the set of all IC and IR mechanisms \( \mu = (q, s) \) that satisfy \( q_i(x) \in [0, \lambda_i] \) (instead of \( q_i(x) \in [0, 1] \)) for every \( x \in \mathbb{R}_+^k \) and \( i = 1, ..., k \). The set of all IC and IR mechanisms, which we denote by \( \mathcal{M} \), is thus the same as \( \mathcal{M}_{1, ..., 1} \).

---

10 The continuity of \( F \) implies that \( X \) cannot be identically zero, and so the optimal revenue \( \text{REV}(X) \) must be positive (just sell the good at a small enough positive price).
Lemma 2 For every $0 < \lambda_1, ..., \lambda_k \leq 1$ we have

$$\text{Rev}(X_1, ..., X_k) = \sup_{\mu \in M_{\lambda_1, ..., \lambda_k}} R(\mu; \tilde{X}_1, ..., \tilde{X}_k),$$

where $\tilde{X}_i := (1/\lambda_i)X_i$ for $i = 1, ..., k$.

Proof. Given $\mu = (q, s)$ with $q_i(x) \in [0, \lambda_i]$ for all $i$, define $\hat{\mu} = (\hat{q}, \hat{s})$ by $\hat{q}_i(x_1, ..., x_k) := (1/\lambda_i)q_i(x_1/\lambda_1, ..., x_k/\lambda_k) \in [0, 1]$ and $\hat{s}(x_1, ..., x_k) := s(x_1/\lambda_1, ..., x_k/\lambda_k)$ (and thus $\hat{b}(x_1, ..., x_k) = b(x_1/\lambda_1, ..., x_k/\lambda_k)$ for the corresponding buyer’s payoff functions). It is immediate to see that $\hat{\mu}$ is IC and IR if and only if $\mu$ is IC and IR, and that $E[s(\tilde{X}_1, ..., \tilde{X}_k)] = E[\hat{s}(X_1, ..., X_k)]$. Conversely, given $\hat{\mu}$ one generates $\mu$ by the reverse transformation. □

3 Overview of the Proof

The first part of the proof is similar to the proofs of Theorems A and B in Hart and Nisan (2017) except that, where they split the buyer’s space of values $(x_1, x_2) \in \mathbb{R}_+^2$ in half along the diagonal $x_1 = x_2$, we split the space into two regions $x_1 \geq \lambda x_2$ and $x_1 < \lambda x_2$ along a possibly nonsymmetric diagonal $x_1 = \lambda x_2$ (the precise value of $\lambda$ will be chosen later). For any two-good mechanism, the revenue in each of the two regions can be estimated by constructing from it appropriate one-good mechanisms, which eventually leads to a key bound: see Proposition 4 in Section 4. (Rather than working directly with the two asymmetric regions, which is cumbersome, the proof simplifies computations by first making an appropriate change of units, which amounts to rescaling the probabilities that the goods are received: see Lemma 2 in Section 2.3.) Once we have the bound given in Proposition 4, we need to estimate the maximum of a certain integral expression—which is essentially the additional revenue that is achievable beyond the separate one-good

11The reader is encouraged to look at these proofs and the explanations there.
revenues—over pairs of nonnegative functions \( \varphi_1, \varphi_2 \) whose sum \( \varphi_1 + \varphi_2 \) is nondecreasing. This is accomplished in Proposition 5 by considering the appropriate extreme functions and then carefully estimating the relevant terms (this is the hardest part of the proof). In Section 6 we put everything together, and, by choosing the best possible \( \lambda \) (specifically, \( \lambda = 1/\sqrt{e} \)), prove the 62% bound (Theorem 7). Then in Section 7 we show the 73% bound for Myerson-regular goods (Theorem 9), and then we also deal with monotonic mechanisms. There are two appendices: Appendix A establishes that, under quite permissive conditions, the seller’s revenue is continuous in the distribution of the buyer’s valuation, a result that we use in our proof, but that is also of independent interest, and Appendix B provides a simple illustration of how the “nonsymmetric diagonal” construct alone can produce useful bounds.

4 Bounding the Revenue by Nonsymmetric Decomposition

This section provides the basic decomposition with respect to a nonsymmetric diagonal (equivalently, we make a corresponding change of units and use the symmetric diagonal; see Section 2.3).

Given a two-good random valuation \((X_1, X_2)\), for \(i = 1, 2\) let \(F_i\) denote the cumulative distribution function of \(X_i\), and let \(f_i, G_i, \) and \(H_i\) be the associated functions as defined in Section 2.2 (namely, the density, tail probability, and cumulative tail probability functions, respectively). We let \(r_i := \text{Rev}(X_i)\) be the optimal revenue that can be obtained from good \(i\), and define two useful auxiliary functions \(K_1\) and \(K_2\):

\[
K_1(t) := f_2(t)(H_1(t) - r_1) - G_1(t)G_2(t), \quad (4)
\]

\[
K_2(t) := f_1(t)(H_2(t) - r_2) - G_1(t)G_2(t). \quad (5)
\]
The following lemma, which slightly generalizes Lemma 19 in Hart and Nisan (2017) (it replaces the factor $1 - q(x_0)$ there with $\lambda - q(x_0)$ here), obtains a better bound on the revenue of a mechanism by “rescaling” its allocation function $q$ so that it covers the entire interval $[0, \lambda]$.

**Lemma 3** Let $X$ be a one-good random valuation with values bounded from below by some $x_0 \geq 0$. Then for every IC mechanism $\mu = (q, s)$ that satisfies $q(x) \leq \lambda$ for all $x \geq x_0$ we have

$$R(\mu; X) \leq (\lambda - q(x_0)) \text{Rev}(X) + s(x_0).$$  \hspace{1cm} (6)

**Proof.** The function $q$ is nondecreasing (because $q$ is the derivative of the buyer’s payoff function $b$, which is convex), and so $q(x_0) \leq q(x) \leq \lambda$ for all $x \geq x_0$. If $q(x_0) = \lambda$ then $q(x) = q(x_0) = \lambda$ for all $x \geq x_0$, hence $s(x) = s(x_0)$ for all $x \geq x_0$ by IC; therefore $E[s(X)] = s(x_0)$ and (6) holds as equality.

If $q(x_0) < \lambda$ then define a new mechanism $\hat{\mu} = (\hat{q}, \hat{s})$ by $\hat{q}(x) := \theta(q(x) - q(x_0))$ and $\hat{s}(x) := \theta(s(x) - s(x_0))$, and thus $\hat{b}(x) := \theta(b(x) - (x - x_0)q(x_0) - b(x_0))$, where $\theta := 1/(\lambda - q(x_0)) > 0$ (so that $0 \leq \hat{q}(x) \leq 1$). It is immediate to verify that $(\hat{q}, \hat{s})$ is an IC and IR mechanism: indeed, $[\hat{q}(x') \cdot x - \hat{s}(x')] - [\hat{q}(x) \cdot x - \hat{s}(x)] = \theta ([q(x') \cdot x - s(x')] - [q(x) \cdot x - s(x)]) \geq 0$, and $\hat{b}(x_0) = 0$. Therefore $\text{Rev}(X) \geq E[\hat{s}(X)] = \theta(E[s(X)] - s(x_0))$, which yields (6). \hfill \blacksquare

We now come to the main result of this section, which generalizes the decomposition of the proofs of Theorems A and B in Hart and Nisan (2017): the revenue from two goods is bounded by the sum of the separate one-good revenues and an additional term (the $K_i$-term), which will be estimated in the next section.

\footnote{It suffices to require $q(x) \leq \lambda$ for $x$ in the support of $X$. As in Hart and Reny (2015), one can always extend a $k$-good mechanism to the whole space $\mathbb{R}_+^k$ without increasing its menu beyond taking closure, and so the bound extends to all $\mathbb{R}_+^k$.}

\footnote{If the values of $X$ are bounded from above by some finite $x_1$, then we can replace $\lambda$ with $q(x_1)$.}
Proposition 4 Let \( X = (X_1, X_2) \) be a two-good random valuation with independent goods (i.e., \( X_1 \) and \( X_2 \) are independent nonnegative real random variables), and let \( \mu = (q, s) \) be a two-good IC, IR, and NPT mechanism that satisfies \( q_i(x) \leq \lambda_i \) for all \( x \in \mathbb{R}_+^2 \) and \( i = 1, 2 \). Then there exist functions \( \varphi_i : \mathbb{R}_+ \to [0, \lambda_i] \) for \( i = 1, 2 \) such that \( \varphi_1 + \varphi_2 \) is a nondecreasing function and
\[
R(\mu; X_1, X_2) \leq \lambda_1 r_1 + \lambda_2 r_2 + \int_0^\infty (\varphi_1(t)K_1(t) + \varphi_2(t)K_2(t))dt; \quad (7)
\]
specifically, \( \varphi_i(t) = q_i(t,t) \).

Proof. The first part of the proof, which yields \( 10 \), goes along the same lines as the proof of Theorem B in Appendix A.1 of Hart and Nisan (2017), but with the appropriate modifications, because here \( X_1 \) and \( X_2 \) are not identically distributed, the mechanism \( \mu \) is not symmetric, and each \( q_i \) is bounded by \( \lambda_i \).

We will write \( Y \) for \( X_1 \) and \( Z \) for \( X_2 \), and so \( X = (Y, Z) \).

For every \( t \geq 0 \) define \( \Phi(t) := b(t,t) \) and \( \varphi_i(t) := q_i(t,t) \). By Proposition \( 1(i) \) the function \( \Phi \) is convex and \( q(t,t) = (\varphi_1(t), \varphi_2(t)) \) is a subgradient of \( b \) at \((t,t)\), and so \( \varphi_1(t) + \varphi_2(t) \) is a subgradient of \( \Phi \) at \( t \). Therefore \( \varphi_1 + \varphi_2 \) is a nondecreasing function, and \( \Phi(u) = \int_0^u (\varphi_1(t) + \varphi_2(t))dt \) (use Corollary 24.2.1 in Rockafellar 1970, recalling that \( \Phi(0) = b(0,0) = 0 \) by NPT).

Consider first the region \( Y \geq Z \). For each fixed value \( z \geq 0 \) of the second good such that \( \mathbb{P}[Y \geq z] > 0 \), define a mechanism \( \mu^z = (q^z, s^z) \) for the first good by replacing the allocation of the second good with an equivalent decrease in payment; that is, the allocation of the first good is unchanged, i.e., \( q^z(y) := q_1(y,z) \), and the payment is \( s^z(y) := s(y,z) - zq_2(y,z) \), for every \( y \geq 0 \); note that the buyer’s payoff remains the same: \( b^z(y) = b(y,z) \). The mechanism \( \mu^z \) is IC and IR for \( y \), since \( \mu \) is IC and IR for \( (y,z) \). Let \( Y^z \) denote the random variable \( Y \) conditional on the event \( Y \geq z \), and consider

\[14\] Notice that \( \Phi \) here is \( 2\Phi \) in Hart and Nisan (2017).
the revenue \( R(\mu^z; Y^z) = \mathbb{E}[s^z(Y^z)] = \mathbb{E}[s^z(Y)|Y \geq z] \) of \( \mu^z \) from \( Y^z \). We have \( Y^z \geq z \), \( q^z(z) = q_1(z, z) = \varphi_1(z) \), and \( s^z(z) = s(z, z) - zq_2(z, z) = zq_1(z, z) - b(z, z) = z\varphi_1(z) - \Phi(z) \), and so, applying Lemma 3 above to \( Y^z \), we have

\[
\mathbb{E}[s^z(Y)|Y \geq z] \leq (\lambda_1 - \varphi_1(z)) \text{REV}(Y^z) + z\varphi_1(z) - \Phi(z) \tag{8}
\]

Since \( \mathbb{P}[Y_t \geq t] = \mathbb{P}[Y \geq t] / \mathbb{P}[Y \geq z] = G_1(t) / \mathbb{P}[Y \geq z] \) for all \( t \geq z \), using (11) we get

\[
\text{REV}(Y^z) = \sup_{t \geq 0} \mathbb{P}[Y^z \geq t] = \sup_{t \geq z} t \cdot \frac{G_1(t)}{\mathbb{P}[Y \geq z]} \leq \sup_{t \geq 0} t \cdot \frac{G_1(t)}{\mathbb{P}[Y \geq z]} = \frac{r_1}{\mathbb{P}[Y \geq z]}
\]

(recall that \( r_1 = \text{REV}(Y) \)). Substitute this into (8), and multiply it by \( \mathbb{P}[Y \geq z] \), to get

\[
\mathbb{E}[s^z(Y)1_{Y \geq z}] \leq (\lambda_1 - \varphi_1(z))r_1 + (z\varphi_1(z) - \Phi(z))\mathbb{P}[Y \geq z]
\]

for all \( z \geq 0 \) (which trivially includes those \( z \) where \( \mathbb{P}[Y \geq z] = 0 \). Taking expectation over the values \( z \) of \( Z \) yields

\[
\mathbb{E}[s^z(Y)1_{Y \geq z}] \leq \lambda_1 r_1 - r_1 \mathbb{E}[\varphi_1(Z)] + \mathbb{E}[(Z\varphi_1(Z) - \Phi(Z))1_{Y \geq z}] \tag{9}
\]

For \( y \geq z \geq 0 \) we have \( s(y, z) = s^z(y) + zq_2(y, z) \leq s^z(y) + zq_2(y, y) = s^z(y) + z\varphi_2(y) \) (by the monotonicity of \( q_2 \) in its second variable, again by the convexity of \( b \), which together with (9) yields

\[
\mathbb{E}[s(Y, Z)1_{Y \geq z}] \leq \mathbb{E}[s^Z(Y)1_{Y \geq z}] + \mathbb{E}[Z\varphi_2(Y)1_{Y \geq z}]
\]

\[
\leq \lambda_1 r_1 - r_1 \mathbb{E}[\varphi_1(Z)] + \mathbb{E}[(Z\varphi_2(Y) + Z\varphi_1(Z) - \Phi(Z))1_{Y \geq z}]
\]

\[
= \lambda_1 r_1 - r_1 \mathbb{E}[\varphi_1(Z)] + \mathbb{E}[(\Lambda\varphi_2(Y) + \Lambda\varphi_1(Z) - \Phi(\Lambda))1_{Y \geq z}],
\]

where we put \( \Lambda := \min\{Y, Z\} \).
Consider next the $Z > Y$. Interchanging $Y$ and $Z$ and using $Z > y$ instead of $Z \geq y$ throughout gives

$$
\mathbb{E}[s(Y, Z)1_{Z>Y}] \leq \lambda_2 r_2 - r_2 \mathbb{E}[\varphi_2(Y)] + \mathbb{E}[(\Lambda \varphi_1(Z) + \Lambda \varphi_2(Y) - \Phi(\Lambda))1_{Z>Y}].
$$

Adding the last two inequalities yields

$$
\mathbb{E}[s(Y, Z)] \leq \lambda_1 r_1 + \lambda_2 r_2 - r_1 \mathbb{E}[\varphi_1(Z)] - r_2 \mathbb{E}[\varphi_2(Y)] + \mathbb{E}[\Lambda \varphi_1(Z) + \Lambda \varphi_2(Y) - \Phi(\Lambda)]
$$

$$
= \lambda_1 r_1 + \lambda_2 r_2 - r_2 \mathbb{E}[\varphi_2(Y)] + \mathbb{E}[\varphi_1(Z) (\Lambda - r_1)] + \mathbb{E}[\varphi_2(Y) (\Lambda - r_2)] - \mathbb{E}[\Phi(\Lambda)]. \quad (10)
$$

Now we have

$$
\mathbb{E}[\varphi_1(Z) (\Lambda - r_1)] = \int_{0}^{\infty} \varphi_1(z)(\mathbb{E}[\min\{Y, z\}] - r_1)f_2(z)dz
$$

$$
= \int_{0}^{\infty} \varphi_1(z)(H_1(z) - r_1)f_2(z)dz \quad (11)
$$

(\mathrm{use} \ \Lambda = \min\{Y, Z\} \ \mathrm{and} \ (12)). \ \mathrm{Similarly},

$$
\mathbb{E}[\varphi_2(Y) (\Lambda - r_2)] = \int_{0}^{\infty} \varphi_2(y)(H_2(y) - r_2)f_1(y)dy. \quad (12)
$$

Let $F_\Lambda$ be the cumulative distribution function of $\Lambda = \min\{Y, Z\}$; then $1 - F_\Lambda(u) = G_\Lambda(u) = G_1(u)G_2(u)$, and

$$
\mathbb{E}[\Phi(\Lambda)] = \int_{0}^{\infty} \Phi(u)dF_\Lambda(u) = -\int_{0}^{\infty} \Phi(u)dG_\Lambda(u)
$$

$$
= [\Phi(u)G_\Lambda(u)]_{0}^{\infty} + \int_{0}^{\infty} \Phi'(u)G_\Lambda(u)du
$$

$$
= \int_{0}^{\infty} \Phi'(u)G_\Lambda(u)du
$$

$$
= \int_{0}^{\infty} (\varphi_1(u) + \varphi_2(u))G_1(u)G_2(u)du, \quad (13)
$$
where we integrated by parts to get the second line\textsuperscript{15} and then used $\Phi(0) = 0$ and $\Phi(\infty)G_{\Lambda}(\infty) = 0$ (because $0 \leq \Phi(u)G_{\Lambda}(u) \leq 2u(r_1/u)(r_2/u) \to 0$ as $u \to \infty$, with $\Phi(u) \leq 2u$ following from $\Phi'(u) \leq 2$).

Substituting (11)–(13) into (10) yields the result. 

5 Bounding the $K_i$-Term

In this section we bound from above the term $\int (\varphi_1K_1 + \varphi_2K_2)$ in (7) over all possible functions $\varphi_i$, which take values in $[0, \lambda_i]$, and whose sum $\varphi_1 + \varphi_2$ is nondecreasing. This term is linear in the $\varphi_i$, and so, if each $\varphi_i$ were nondecreasing, it would suffice to consider only the extreme functions that take the values 0 and $\lambda_i$ (because any nondecreasing function is an average of such functions; see the remark below). However, we only require the sum to be nondecreasing, which requires a more delicate analysis; see Proposition 5. This result is then applied to our specific functions $K_1$ and $K_2$ to get the bound in Proposition 6 (this constitutes the core of the proof).

From now on we will assume without loss of generality that $\lambda_1 \leq \lambda_2$, and so $0 < \lambda_1 \leq \lambda_2 \leq 1$. Let $K_1, K_2 : \mathbb{R}_+ \to \mathbb{R}$ be two functions, and define

$$I := \sup_{\varphi_1, \varphi_2} \int_0^\infty (\varphi_1(t)K_1(t) + \varphi_2(t)K_2(t)) \, dt,$$

where the supremum is taken over all functions $\varphi_i : \mathbb{R}_+ \to [0, \lambda_i]$ such that $\varphi := \varphi_1 + \varphi_2$ is a nondecreasing function.

\textsuperscript{15}Formally, we integrate by parts on a finite interval $[0, M]$ and then let $M \to \infty$. The functions $G_{\Lambda}$ and $\Phi$ are absolutely continuous (because $G_i = 1 - F_i$ for $i = 1, 2$ are absolutely continuous and $G_{\Lambda} = G_1G_2$, and $\Phi$ is convex and continuous).
To estimate $I$, for any $0 \leq a \leq b \leq c \leq \infty$ define

$$I(a, b, c) := \int_a^b \lambda_1 \max\{K_1(t), K_2(t)\} \, dt$$

$$\quad + \int_b^c (\lambda_2 - \lambda_1)K_2(t) + \lambda_1 \max\{K_1(t), K_2(t)\} \, dt$$

$$\quad + \int_c^\infty (\lambda_1 K_1(t) + \lambda_2 K_2(t)) \, dt$$

$$= \lambda_1 \int_a^c \max\{K_1(t), K_2(t)\} \, dt + \lambda_1 \int_c^\infty (K_1(t) + K_2(t)) \, dt$$

$$\quad + (\lambda_2 - \lambda_1) \int_b^\infty K_2(t) \, dt. \quad (14)$$

It is immediate to see that $I(a, b, c)$ is nothing other than $\int (\varphi_1 K_1 + \varphi_2 K_2)$ for the following functions $\varphi_1$ and $\varphi_2$:

|               | $a \leq t < b$ | $b \leq t < c$ | $t \geq c$ |
|---------------|----------------|----------------|------------|
| $\varphi_1(t)$| $\lambda_1 1_{K_1(t) \geq K_2(t)}$ | $\lambda_1 1_{K_1(t) \geq K_2(t)}$ | $\lambda_1$ |
| $\varphi_2(t)$| $\lambda_1 1_{K_1(t) < K_2(t)}$ | $\lambda_1 1_{K_1(t) < K_2(t)} + \lambda_2 - \lambda_1$ | $\lambda_2$ |

Their sum $\varphi_1 + \varphi_2$ then equals

|               | $a \leq t < b$ | $b \leq t < c$ | $t \geq c$ |
|---------------|----------------|----------------|------------|
| $\varphi_1(t) + \varphi_2(t)$| $\lambda_1$ | $\lambda_2$ | $\lambda_1 + \lambda_2$ |

which is a nondecreasing function, and so $I(a, b, c) \leq I$.

**Proposition 5** Let $0 < \lambda_1 \leq \lambda_2 \leq 1$. Then

$$I = \sup_{0 \leq a \leq b \leq c \leq \infty} I(a, b, c).$$
Remark. We will use the following well-known result. Every nondecreasing function $\psi : [u, v] \to [0, 1]$ (where $-\infty \leq u \leq v \leq \infty$) can be expressed as an (integral) average of nondecreasing functions that take only the values $0$ and $1$. More generally, every nondecreasing function $\psi : [u, v] \to [\alpha, \beta]$ (where $\alpha \leq \beta$ are finite) can be expressed as an average of nondecreasing functions that take only the two values $\alpha$ and $\beta$ (when $\alpha < \beta$, apply the above to $(\psi - \alpha)/(\beta - \alpha)$, which takes values in $[0, 1]$). Therefore, when we maximize a linear functional $\int_u^v \psi(t)K(t)dt$ over all nondecreasing functions $\psi : [u, v] \to [\alpha, \beta]$, it suffices to consider those functions that take only the two extreme values $\alpha$ and $\beta$.

Proof. We have seen above that $I \geq I(a, b, c)$ for every $a, b, c$. We now show that the supremum in $I$ cannot be higher.

For each $t$, given $\varphi(t) = \varphi_1(t) + \varphi_2(t)$, the expression $\varphi_1(t)K_1(t) + \varphi_2(t)K_2(t)$ is maximized by putting as much weight as possible—subject to the constraints $0 \leq \varphi(t) \leq \lambda_i$—on the higher of $K_1(t)$ and $K_2(t)$. This gives the following upper bounds on $\varphi_1(t)K_1(t) + \varphi_2(t)K_2(t)$:

- $\varphi(t) \max\{K_1(t), K_2(t)\}$ for every $t$ in the interval where $0 \leq \varphi(t) \leq \lambda_1$;
- $(\varphi(t) - \lambda_1)K_2(t) + \lambda_1 \max\{K_1(t), K_2(t)\}$ for every $t$ in the interval where $\lambda_1 \leq \varphi(t) \leq \lambda_2$ (because $\varphi_1(t) \leq \lambda_1$ implies $\varphi_2(t) \geq \varphi(t) - \lambda_1$); and
- $(\varphi(t) - \lambda_2)K_1(t) + (\varphi(t) - \lambda_1)K_2(t) + (\lambda_1 + \lambda_2 - \varphi(t)) \max\{K_1(t), K_2(t)\}$ for every $t$ in the interval where $\lambda_2 \leq \varphi(t) \leq \lambda_1 + \lambda_2$.

\[16\text{Assume first that } \psi(v) = 1. \text{ If } \psi \text{ is a right-continuous function then } \psi \text{ may be viewed as a cumulative distribution function on } [u, v], \text{ and we have } \psi(t) = \int_{[u,v]} d\psi(x) = \int_{[u,v]} 1_{[x,v]}(t) d\psi(x) \text{ for every } t \in [u, v] \text{ (where } 1_E \text{ is the indicator function of the set } E, \text{ i.e., } 1_E(t) = 1 \text{ if } t \in E \text{ and } 1_E(t) = 0 \text{ otherwise). If } \psi \text{ is not necessarily right-continuous, let } \psi_+(t) := \lim_{t \uparrow} \psi(t) \text{ (which is right-continuous)}, \psi_-(t) := \lim_{t \downarrow} \psi(t), \text{ and take } \lambda_i \in [0, 1] \text{ such that } \psi(t) = \lambda_i \psi_+(t) + (1 - \lambda_i) \psi_-(t); \text{ then } \psi = \int_{[u,v]} (\lambda_x 1_{[x,v]} + (1 - \lambda_x) 1_{[x,v]}) d\psi_+(x). \]

If $0 < \psi(v) < 1$ then $\psi = \psi(v) \tilde{\psi} + (1 - \psi(v))0$, where $\tilde{\psi}(t) := \psi(t)/\psi(v)$ and $0$ is the zero function (i.e., $0(t) = 0$ for all $t$), and we apply the above to $\tilde{\psi}$. Finally, if $\psi(v) = 0$ then $\psi = 0$. 

18
In each one of these three intervals the bound is affine in $\varphi$ and so, by the remark above, when maximizing over nondecreasing $\varphi$, it suffices to consider solely those functions $\varphi$ that take only the corresponding two extreme values. Altogether, such a $\varphi$ takes only the values $0, \lambda_1, \lambda_2$, and $\lambda_1 + \lambda_2$, say on the intervals $(0, a), (a, b), (b, c)$, and $(c, \infty)$, respectively—and then $\int (\varphi_1 K_1 + \varphi_2 K_2)$ becomes precisely $I(a, b, c)$. Thus indeed $I \leq \sup I(a, b, c)$. ■

We now come to the main argument of our proof, which yields, using Proposition 5, an upper bound on the $K_i$-term for our specific functions $K_i$.

**Proposition 6** Let $0 < \lambda_1 \leq \lambda_2 \leq 1$, and let $K_1, K_2$ be given by (2) and (3). Then

$$I \leq \frac{1}{e} \left( \lambda_2 r_1 + \lambda_1 r_2 + \lambda_1 (e - 1) \min \{r_1, r_2\} \right).$$

**Proof.** Recalling (2), we have the following: for each $i = 1, 2$, the function $H_i(t)$ is continuous and strictly increasing at each $t$ in the support of $X_i$ (because $G_i(t) > 0$ there), and $H_i(\infty) = E[X_i] \geq r_i$ (with strict inequality unless $X_i$ is constant, in which case everything trivializes). Therefore there exists a finite $\tau_i$ such that $H_i(\tau_i) = r_i$; since for all $t < r_i$ we have $H_i(t) \leq t < r_i$ (because $G_i \leq 1$), it follows that

$$\tau_i \geq r_i.$$

Put $L_i(t) := G_j(t)(H_i(t) - r_i)$; taking derivatives gives

$$L'_i(t) = -f_j(t)(H_i(t) - r_i) + G_j(t)G_i(t) = -K_i(t).$$

We will use the following estimates:

$$\int_u^\infty G_1(t)G_2(t)dt \leq \int_u^\infty \frac{r_1 r_2}{t} \frac{dt}{t} = \frac{r_1 r_2}{u}$$

(15)
for every $u > 0$ (because $G_i(t) \leq r_i/t$);

$$H_i(u) - r_i \leq r_i \log \frac{u}{r_i}$$

for every $u \geq r_i$ (recall (3)); and, thus,

$$L_i(u) = G_j(u)(H_i(u) - r_i) \leq \frac{r_i r_j}{u} \log \frac{u}{r_i}$$  \hspace{1cm} (16)

for every $u \geq r_i$. The last inequality implies that $L_i(u) \to 0$ as $u \to \infty$, and so

$$\int_u^\infty K_i(t)dt = [-L_i(t)]_u^\infty = L_i(u).$$  \hspace{1cm} (17)

Finally, letting $\{i,j\} = \{1,2\}$, we have

$$L_i(u) \leq \frac{1}{e} r_j \text{ and } L_i(u) + \frac{r_i r_j}{u} \leq r_j$$  \hspace{1cm} (18) and (19)

for every $u \geq r_i$ (use (16) together with $\log x/x \leq 1/e$ and $(\log x + 1)/x \leq 1$ for all $x > 0$; note that there is no typo here: these bounds on $L_i$ use $r_j$ rather than $r_i$).

We need to bound $I(a,b,c)$. For the last term of (14) we have, by (17) and (18),

$$\int_b^\infty K_2(t)dt \leq \frac{1}{e} r_1,$$  \hspace{1cm} (20)

and so it remains to estimate $J(a,c) := \int_a^c \max\{K_1,K_2\} + \int_c^\infty (K_1 + K_2)$.

A main difficulty in doing so is that the $K_i$ are neither nonnegative nor monotonic, and may change signs many times. To handle this we define for each $i$ an auxiliary function $M_i(t) := f_j(t)(H_i(t) - r_i) = K_i(t) + G_1(t)G_2(t)$, which vanishes at $t = \tau_i$, is nonpositive before $\tau_i$, and nonnegative after $\tau_i$; i.e., $M_i(t) \geq 0$ for $t \geq \tau_i$ and $M_i(t) \leq 0$ for $t \leq \tau_i$.

We distinguish three cases according to the location of $a$ relative to $\tau_1$ and
\(\tau_2\) (the points where \(M_1\) and \(M_2\) change sign); without loss of generality assume that \(\tau_1 \leq \tau_2\).

- **Case 1.** \(a \geq \max\{\tau_1, \tau_2\} = \tau_2\). For every \(t \geq a\) we have \(M_i(t) \geq 0\) (because \(t \geq a \geq \tau_i\)), and thus

  \[
  \max\{K_1(t), K_2(t)\} = \max\{M_1(t), M_2(t)\} - G_1(t)G_2(t) \\
  \leq M_1(t) + M_2(t) - G_1(t)G_2(t) \\
  = K_1(t) + K_2(t) + G_1(t)G_2(t).
  \]

Since we clearly also have \(K_1 + K_2 \leq K_1 + K_2 + G_1G_2\), we get

\[
J(a, c) \leq \int_a^{\infty} (K_1(t) + K_2(t) + G_1(t)G_2(t))dt \\
\leq L_1(a) + L_2(a) + \frac{r_1r_2}{a} =: \bar{J}(a)
\]

by (17) and (15). If, say, \(r_k \leq r_\ell\) (where \(\{k, \ell\} = \{1, 2\}\)) then using (18) for \(i = k\) and (19) for \(i = \ell\) (recall that \(a \geq \tau_i \geq r_i\) for both \(i\)) yields

\[
J(a, c) \leq \bar{J}(a) \leq \frac{1}{e}r_\ell + r_k.
\]

- **Case 2.** \(\tau_1 \leq a < \tau_2\). In the range \(t \in [a, \tau_2) \subseteq [\tau_1, \tau_2)\) we have \(M_1(t) \geq 0 \geq M_2(t)\), and so \(K_1(t) \geq K_2(t)\) and \(K_2(t) \leq 0\); therefore both \(\max\{K_1(t), K_2(t)\}\) and \(K_1(t) + K_2(t)\) are \(\leq K_1(t)\), and thus, regardless of

\[\text{\footnotesize{\textsuperscript{17}}}\text{The expression } J(a, b) \text{ that we estimate now is symmetric in } i = 1, 2, \text{ and so the assumption that } \lambda_1 \leq \lambda_2 \text{ is irrelevant here; we thus assume that } \tau_1 \leq \tau_2.\]
\[\text{\footnotesize{\textsuperscript{18}}}\text{This is the inequality } \max\{x, y\} \leq x + y + z \text{ whenever } x, y \geq -z.\]
\[\text{\footnotesize{\textsuperscript{19}}}\text{A slightly better estimate of } (2/\sqrt{e})\sqrt{r_1r_2} \text{ may be obtained here by directly maximizing } \bar{J}(a) \text{ over } a; \text{ however, this will not improve the overall estimate, due to Cases 2 and 3.}\]
where $c$ is,

$$J(a, c) \leq \int_a^{\tau_2} K_1(t) dt + \bar{J}(\tau_2) = (L_1(a) - L_1(\tau_2)) + \left( L_1(\tau_2) + L_2(\tau_2) + \frac{r_1 r_2}{\tau_2} \right)$$

$$= L_1(a) + \frac{r_1 r_2}{\tau_2} \leq \frac{1}{e} r_2 + \min\{r_1, r_2\} \leq \frac{1}{e} r_\ell + r_k,$$

where we have used: $L_2(\tau_2) = 0$ and $\tau_2 = \max\{\tau_1, \tau_2\} \geq \max\{r_1, r_2\}$ (because $\tau_i \geq r_i$).

• **Case 3.** $a < \min\{\tau_1, \tau_2\} = \tau_1$. For every $t \leq \tau_1$ we have $K_1(t) \leq M_1(t) \leq 0$ and $K_2(t) \leq M_2(t) \leq 0$, and so both $\max\{K_1(t), K_2(t)\}$ and $K_1(t) + K_2(t)$ are $\leq 0$ in the interval $[a, \tau_1]$. Therefore $J(a, c) \leq J(\tau_1, \max\{c, \tau_1\})$, to which we apply Case 2 (with $a = \tau_1$).

Thus in all three cases the bound on $J(a, c)$ is $(1/e) r_\ell + r_k = (1/e)(r_1 + r_2) + (1 - 1/e) \min\{r_1, r_2\}$; together with (20), we get

$$I(a, b, c) \leq \frac{\lambda_1}{e} r_1 + \frac{\lambda_1}{e} r_2 + \frac{\lambda_1(e - 1)}{e} \min\{r_1, r_2\} + \frac{\lambda_2 - \lambda_1}{e} r_1,$$

completing the proof. ■

**Remark.** If $\varphi_1$ and $\varphi_2$ are each required to be nondecreasing (rather than just their sum), then we get a smaller bound on $\int (\varphi_1 K_1 + \varphi_2 K_2)$, namely:

$$\sup_{\varphi_1, \varphi_2} \int_0^{\infty} (\varphi_1(t) K_1(t) + \varphi_2(t) K_2(t)) dt$$

$$= \sup_{0 \leq a \leq \infty} \lambda_1 \int_a^{\infty} K_1(t) dt + \sup_{0 \leq b \leq \infty} \lambda_2 \int_b^{\infty} K_2(t) dt \leq \frac{\lambda_1}{e} r_2 + \frac{\lambda_2}{e} r_1$$

(see the remark preceding the proof of Proposition 5 together with (17) and (18)). Therefore, for mechanisms $\mu = (q, s)$ where $q_1(t, t)$ and $q_2(t, t)$ are monotonic—such as, for instance, symmetric mechanisms, where $\Box \Box q_1(t, t) = \Box$,

---

20This proves Theorem B of Hart and Nisan (2017) for two independent and identically distributed goods.
$q_2(t, t)$—we get, taking $\lambda_1 = \lambda_2 = 1$ in Proposition 4,

$$R(\mu; X) \leq r_1 + r_2 + \frac{1}{e} r_1 + \frac{1}{e} r_2 = \left(1 + \frac{1}{e}\right)(r_1 + r_2).$$

This yields the bound $e/(e + 1)$, which is better than $\sqrt{e}/(\sqrt{e} + 1)$.

6 Completing the Proof

Combining the results of the previous two sections yields the first part of our Main Result:

**Theorem 7** Let $X = (X_1, X_2)$ be a two-good random valuation with independent goods. Then

$$\frac{\text{SRev}(X_1, X_2)}{\text{Rev}(X_1, X_2)} \geq \frac{\sqrt{e}}{\sqrt{e} + 1} \approx 0.62.$$ 

**Proof.** Let $R_i := \text{Rev}(X_i)$; thus $\text{SRev}(X_1, X_2) = R_1 + R_2$. Given $0 < \lambda_1 \leq \lambda_2$, put $\tilde{X}_i := X_i/\lambda_i$ and $r_i := \text{Rev}(\tilde{X}_i) = R_i/\lambda_i$. Using Lemma 2, Proposition 4 for $(\tilde{X}_1, \tilde{X}_2)$, and then Proposition 6 yields

$$\text{Rev}(X_1, X_2) \leq \lambda_1 \frac{R_1}{\lambda_1} + \lambda_2 \frac{R_2}{\lambda_2} + \frac{\lambda_1 R_1}{e \lambda_1} + \frac{\lambda_2 R_2}{e \lambda_2} + \frac{\lambda_1(e - 1)}{e} \min \left\{ \frac{R_1}{\lambda_1}, \frac{R_2}{\lambda_2} \right\}$$

$$\leq R_1 + R_2 + \frac{1}{\lambda e} R_1 + \frac{\lambda}{e} R_2 + \frac{\lambda(e - 1)}{e} R_2$$

$$= R_1 + R_2 + \frac{1}{\lambda e} R_1 + \lambda R_2,$$

where in the second line we put $\lambda := \lambda_1/\lambda_2 \in (0, 1]$ and used $\min \{R_1/\lambda_1, R_2/\lambda_2\} \leq R_2/\lambda_2$. The final expression equals $(1 + 1/\sqrt{e})(R_1 + R_2)$ when $\lambda = 1/\sqrt{e}$, completing the proof.\footnote{The results of the previous sections will be applied to the rescaled $\tilde{X}_i = X_i/\lambda_i$, and so we will use $r_i$ for the revenue of $\tilde{X}_i$, and $R_i$ for the revenue of the original $X_i$.}

\footnote{One may check that $1 + 1/\sqrt{e}$ is the best bound that is independent of $R_1$ and $R_2$.}
7 Regular Goods and Monotonic Mechanisms

In this section we prove the second part of our Main Result, namely, the better bound of 73% for regular goods (and also for monotonic mechanisms). We will use here only the symmetric diagonal decomposition (i.e., $\lambda_1 = \lambda_2 = 1$).

Following Myerson (1981), we say that a one-dimensional random variable $X$ is weakly regular if its support is an interval $[\alpha, \beta]$ with $0 \leq \alpha < \beta \leq \infty$, on which it has a density function $f(t)$ that is positive and continuous, and the resulting “virtual valuation function” $t - G(t)/f(t)$ is nondecreasing (Myerson’s regularity condition requires the virtual valuation to be strictly increasing).

**Lemma 8** Assume that $X_1$ and $X_2$ are weakly regular. Then $K_i(u) > 0$ implies that $K_i(v) \geq 0$ for all $v > u$, for $i = 1, 2$.

**Proof.** Let $[\alpha_i, \beta_i]$ be the support of $X_i$. Assume by way of contradiction that, say, $K_1(u) > 0$ and $K_1(v) < 0$ for some $v > u$. First, $K_1(u) > 0$ implies that $f_2(u) > 0$ and $H_1(u) - r_1 > 0$ (otherwise $K_1(u) \leq - G_1(u) G_2(u) \leq 0$), and so $\alpha_2 \leq u \leq \beta_2$ and $u > \alpha_1$ (because $H_1$ is nondecreasing and $H_1(\alpha_1) \leq \alpha_1 = \alpha_1 \cdot G_1(\alpha_1) \leq r_1$). Second, $K_1(v) < 0$ implies that $G_1(v) > 0$ and $G_2(v) > 0$ (otherwise $K_1(v) = f_2(v)(H_1(v) - r_1) \geq f_2(v)(H_1(u) - r_1) \geq 0$ since $H_1$ is nondecreasing), and so $v < \beta_1$ and $v < \beta_2$ (because $G_i(\beta_i) = 0$).

Together with $u < v$ it follows that $u$ and $v$ both lie in the interval $[\alpha_1, \beta_1)$ where $f_2(t) > 0$, $G_1(t) > 0$, and $H_1(t) - r_1 > 0$. But in that interval the function $\kappa$, defined by

$$\kappa(t) := \frac{K_1(t)}{f_2(t) G_1(t)} = \left( \frac{H_1(t) - r_1}{G_1(t)} - t \right) + \left( t - \frac{G_2(t)}{f_2(t)} \right),$$

(when $R_1 = R_2$ the above expression is minimized only at $\lambda = 1/\sqrt{e}$).

Notice that we allow $\beta = \infty$, in which case the interval is understood to be $[\alpha, \infty)$.

I.e., $t_1 < u < v < \beta_1$, where $t_1 > \alpha_1$ is the point where $H_1(t_1) = r_1$, and $\alpha_2 \leq u < v < \beta_2$. 

24
is increasing—the derivative of the first term is $f_1(t)(H_1(t) - r_1)/G_1^2(t) > 0$, and the second term is nondecreasing by regularity. Therefore we cannot have $\kappa(u) > 0$ and $\kappa(v) < 0$, which contradicts the assumption that $K_1(u) > 0$ and $K_1(v) < 0$. ■

This yields the second part of our Main Result:

**Theorem 9** Let $X = (X_1, X_2)$ be a two-good random valuation with independent and weakly regular goods. Then

$$
\frac{\text{SRev}(X_1, X_2)}{\text{Rev}(X_1, X_2)} \geq \frac{e}{e + 1} \approx 0.73.
$$

**Proof.** Take $\lambda_1 = \lambda_2 = 1$ and let $r_i = \text{Rev}(X_i)$. Lemma 8 implies that if $K_i(t)$ is positive anywhere then it is nonnegative from that point on, and so either (i) there is some finite $u \geq 0$ such that $K_i(t) \leq 0$ for $t < u$ and $K_i(t) \geq 0$ for $t \geq u$ or (ii) $K_i(t) \leq 0$ for all $t \geq 0$. Therefore, for any function $\varphi_i$ with values in $[0, 1]$, we have

$$
\int_0^\infty \varphi_i(t)K_i(t)dt \leq \int_u^\infty K_i(t)dt = L_i(u) \leq \frac{1}{e} r_j
$$

in case (i) (by (17) and (18)), and $\int_0^\infty \varphi_i(t)K_i(t)dt \leq 0$ in case (ii). Altogether $\int \varphi_1K_1 + \int \varphi_2K_2 \leq (1/e)r_2 + (1/e)r_1$, and so Proposition 4 gives $\text{Rev}(X) \leq (1 + 1/e)(r_1 + r_2)$, proving the result. ■

Next, let $\text{MonRev}(X)$ denote the maximal revenue that can be obtained using monotonic mechanisms, i.e., mechanisms $\mu = (q, s)$ for which the function $s(x)$ is nondecreasing in $x$.

**Proposition 10** Let $X = (X_1, X_2)$ be a two-good random valuation with independent and weakly regular goods. Then

$$
\frac{\text{SRev}(X_1, X_2)}{\text{MonRev}(X_1, X_2)} \geq \frac{e}{e + 1} \approx 0.73.
$$

25
Proof. Put $r_i := \text{Rev}(X_i)$, and let $V_i$ be the “equal-revenue” (ER) random valuation with the same revenue $r_i$ as $X_i$; i.e., its tail distribution function is $\hat{G}_i(t) = \min\{r_i/t, 1\} \geq G_i(t)$. Take $V_1$ and $V_2$ to be independent, and put $V = (V_1, V_2)$. Because $V_i$ first-order stochastically dominates $X_i$, for every monotonic mechanism $\mu = (q, s)$ we have $R(\mu; X) = \mathbb{E}[s(X_1, X_2)] \leq \mathbb{E}[s(V_1, V_2)] = R(\mu; V)$. Therefore

$$\text{MonRev}(X) \leq \text{MonRev}(V) \leq \text{Rev}(V).$$

The ER-good $V_i$ is weakly regular (because on its support $[r_i, \infty)$ the virtual valuation function $t - \hat{G}_i(t)/\hat{f}_i(t)$ is identically 0), and so $\text{SRev}(V) \geq e/(e + 1)\text{Rev}(V)$ by Theorem 9 together with $\text{SRev}(X) = \text{SRev}(V)$ (by construction) and $\text{MonRev}(X) \leq \text{Rev}(V)$ (see above), the result follows.

A Appendix: Revenue Continuity

This appendix deals with the continuity of the revenue with respect to valuations, which is of independent interest. Take a sequence of $k$-good valuations $X^n$ that converges in distribution to the $k$-good random valuation $X$; does the sequence of revenues $\text{Rev}(X^n)$ converge to $\text{Rev}(X)$?

Even in the one-good case this need not be so: for each $n$ let $X^n$ be the one-good valuation that takes value 0 with probability $1 - 1/n$ and value $n$ with probability $1/n$. Then $X^n$ converges in distribution to the valuation $X$ that takes value 0 with probability 1. But $\text{Rev}(X^n) = 1$ (with the posted

\[25\]Only the distribution of a random valuation $X$ matters for the revenue achievable from $X$; it is thus natural to consider what happens when $X^n$ converges in distribution to $X$. Formally, convergence in distribution is equivalent to the cumulative distribution functions converging pointwise at all points of continuity of the limit cumulative distribution. Informally, being close in distribution means that the probabilities of nearby values are close (see, for instance, \[22\] below). Billingsley’s (1968) book is a good reference for the concepts used here.
price of \( n \) while \( \text{Rev}(X) = 0 \).

We will show that if the valuations all lie in a bounded set—more generally, if the random valuations are uniformly integrable—then the limit of the revenues equals the revenue of the limit. We emphasize that all the results in this appendix are for \emph{general} \( k \)-good valuations for any \( k \geq 1 \), whether the goods are independent or not.\footnote{Monteiro (2015) establishes continuity of the optimal revenue in the one-good case with \( n \) independent buyers, when the valuations are bounded and the limit distributions are continuous (his proof uses the characterization of the optimal mechanism).}

Some notation. First, it will be convenient to work here with the \( \ell_1 \)-norm on \( \mathbb{R}^k \), i.e., \( ||x||_1 = \sum_{i=1}^{k} |x_i| \). The \( \ell_1 \)-norm of a valuation \( x \) in \( \mathbb{R}_+^k \) provides a simple bound on the seller’s payoff in any mechanism \( \mu = (q, s) \) that is IR: \( s(x) \leq q(x) \cdot x \leq \sum_i x_i = ||x||_1 \); thus, if a random valuation \( X \) satisfies \( ||X||_1 \leq M \) then \( \text{Rev}(X) \leq M \). Second, the \emph{Prohorov distance} between the distributions of \( X \) and \( Y \), which we denote by \( \text{Dist}(X,Y) \), is defined as the infimum of all \( \rho > 0 \) such that

\[
\mathbb{P}[X \in A] \leq \mathbb{P}[Y \in B_\rho(A)] + \rho, \quad \text{and} \quad (22) \\
\mathbb{P}[Y \in A] \leq \mathbb{P}[X \in B_\rho(A)] + \rho
\]

for all measurable sets \( A \), where \( B_\rho(A) := \{ y : ||y - x||_1 < \rho \text{ for some } x \in A \} \) is the \( \rho \)-neighborhood of \( A \). Thus \( 0 \leq \text{Dist}(X,Y) \leq 1 \), and \( X^n \) converges in distribution to \( X \), which we write as \( X^n \xrightarrow{D} X \), if and only if \( \text{Dist}(X,X^n) \to 0 \) \footnote{This affects only the various constants below.} (again, see Billingsley 1968).

The basic result is that in the bounded case the distance between the revenues of two random valuations is uniformly bounded by a function of the Prohorov distance between their distributions.

\textbf{Proposition 11} Let \( X \) and \( Y \) be \( k \)-good valuations with bounded values, say,
\[ ||X||_1, ||Y||_1 \leq M \text{ for some } M \geq 1. \]

Theorem

\[ |\text{Rev}(X) - \text{Rev}(Y)| \leq (2M + 1)\sqrt{\text{Dist}(X, Y)}. \]

**Proof.** If \( \text{Dist}(X, Y) = 1 \) there is nothing to prove, since both revenues are between 0 and \( M \).

Let thus \( 0 < \rho < 1 \) be such that (22) holds for every measurable set \( A \subseteq \mathbb{R}_+^k \), and take \( \alpha \) so that \( \rho \leq \alpha < 1 \) (the value of \( \alpha \) will be determined later). Denote by \( D_M := \{ x \in \mathbb{R}_+^k : ||x||_1 \leq M \} \) the domain of values of \( X \) and \( Y \).

Let \( \mu = (q, s) \) be an IC, IR, and NPT mechanism, and let \( b \) be its buyer payoff function. We define a new mechanism \( \tilde{\mu} \) by lowering all payments by a factor of \( 1 - \alpha \) (and letting the buyer reoptimize). Thus, let \( \text{cl } W \subset \mathbb{R}_+^{k+1} \) be the closure of the set \( W := \{ (q(x), (1-\alpha)s(x)) : x \in D_M \} \). For each \( x \in D_M \) let \( (\tilde{q}(x), \tilde{s}(x)) \) be a maximizer for (29) \( \tilde{b}(x) := \max_{(g,t) \in \text{cl } W} g \cdot x - t \). Then the mechanism \( \tilde{\mu} = (\tilde{q}, \tilde{s}) \) is IC (by the maximizer definition), IR (because \( \tilde{b}(x) \geq b(x) + \alpha s(x) \), which is nonegative since \( \mu \) is IR and NPT), and NPT (because \( \mu \) is NPT). \( \tilde{b}(x) \geq b(x) + \alpha s(x) \), which is nonegative since \( \mu \) is IR and NPT).

Let \( x, y \in D_M \) be such that \( ||x-y||_1 \leq \rho \). Then \( (\tilde{q}(y), \tilde{s}(y)) \in \text{cl } W \) can be approximated by elements of \( W \): for every \( \epsilon > 0 \) there is \( z \in D_M \) such that, in particular, \( |\tilde{s}(y) - (1-\alpha)s(z)| \leq \epsilon \) and \( |[\tilde{q}(y) \cdot y - \tilde{s}(y)] - [q(z) \cdot y - (1-\alpha)s(z)]| \leq \epsilon \).

---

28 We have not attempted to optimize the bound here.

29 This maximum is attained because \( W \) is bounded, namely, \( W \subseteq [0,1]^k \times [0,M] \), and so \( \text{cl } W \) is a compact set.

30 Hart and Reny (2015) use this device of applying a small uniform discount to the buyer’s payments to show that, at an arbitrarily small cost, the seller can perturb any IC and IR mechanism so that the buyer breaks any indifference in the seller’s favor (the resulting mechanism is thus called *seller-favorable*).
\( \varepsilon \). We then have

\[
q(z) \cdot y - (1 - \varepsilon)s(z) + \varepsilon \geq \tilde{q}(y) \cdot y - \tilde{s}(y)
\]

\[
\geq q(x) \cdot y - (1 - \alpha)s(x)
\]

\[
= q(x) \cdot x - s(x) + q(x) \cdot (y - x) + \alpha s(x)
\]

\[
\geq q(z) \cdot x - s(z) + q(x) \cdot (y - x) + \alpha s(x),
\]

where the second inequality follows because \((q(x), (1 - \alpha)s(x)) \in W\), and the last inequality follows because \((q, s)\) is IC. Rearranging gives

\[
\alpha (s(z) - s(x)) \geq (q(x) - q(z)) \cdot (y - x) - \varepsilon.
\]

Now \(|(q(x) - q(z)) \cdot (y - x)| \leq \rho\) (because \(q(x), q(z) \in [0, 1]^k\) and \(||y - x||_1 \leq \rho\)), and so

\[
(1 - \alpha)s(z) \geq (1 - \alpha)s(x) - \frac{1 - \alpha}{\alpha}(\rho + \varepsilon).
\]

Recalling that \(\tilde{s}(y) \geq (1 - \alpha)s(z) - \varepsilon\) yields

\[
\tilde{s}(y) \geq (1 - \alpha)s(x) - \frac{1 - \alpha}{\alpha}(\rho + \varepsilon) - \varepsilon;
\]

as \(\varepsilon > 0\) was arbitrary, we have

\[
\tilde{s}(y) \geq (1 - \alpha)s(x) - \frac{1 - \alpha}{\alpha}\rho,
\]

and so, using \(s(x) \leq ||x||_1 \leq M\),

\[
\tilde{s}(y) \geq s(x) - \beta,
\]

where \(\beta := \alpha M + (1 - \alpha)\rho/\alpha\).

For each \(t > 0\) put \(A(t) := \{x \in D_M : s(x) \geq t\}\) and \(\tilde{A}(t) := \{x \in D_M : \tilde{s}(x) \geq t\}\). Inequality (23), which applies whenever \(||y - x||_1 \leq \rho\), implies
that $\tilde{A}(t - \beta) \supseteq B_\rho(A(t))$, and so

$$\text{Rev}(Y) \geq \mathbb{E}[\tilde{s}(Y)] = \int_0^\infty \mathbb{P}[Y \in \tilde{A}(t)] \, dt \geq \int_0^M \mathbb{P}[Y \in \tilde{A}(t - \beta)] \, dt \geq \int_\beta^M \mathbb{P}[Y \in B_\rho(A(t))] \, dt \geq \int_\beta^M \mathbb{P}[X \in A(t)] \, dt - (M - \beta)\rho \geq \int_0^M \mathbb{P}[X \in A(t)] \, dt - \beta - (M - \beta)\rho = \mathbb{E}[s(X)] - \beta',$$

where $\beta' := \beta + (M - \beta)\rho$ (for the fourth inequality we have used (22) and $\beta \leq M$, which follows from $\rho \leq \alpha$ and $M \geq 1$).

Taking $\alpha = \sqrt{\rho}$ gives $\beta \leq (M + 1)\sqrt{\rho}$ and $\beta' \leq \beta + M\rho \leq (2M + 1)\sqrt{\rho}$. The displayed inequality above holds for every $\mu$ and every $\rho > \text{Dist}(X,Y)$, and so $\text{Rev}(Y) \geq \text{Rev}(X) - (2M + 1)\sqrt{\text{Dist}(X,Y)}$. Interchanging $X$ and $Y$ completes the proof.

A sequence of random variables $(X^n)_{n \geq 1}$ is uniformly integrable if for every $\varepsilon > 0$ there is a finite $M$ such that $\mathbb{E}[||X^n||_1 1_{||X^n||_1 > M}] \leq \varepsilon$ for all $n$.

**Theorem 12** Let $X^n$ be a sequence of $k$-good random valuations that converges in distribution to the $k$-good random valuation $X$. Then

$$\liminf_{n \to \infty} \text{Rev}(X^n) \geq \text{Rev}(X).$$

Moreover, if the sequence $X^n$ is uniformly integrable, then

$$\lim_{n \to \infty} \text{Rev}(X^n) = \text{Rev}(X) < \infty.$$

**Proof.** For every $M > 0$ and every $k$-good valuation $X$, denote $X_{(M)} := X1_{||X||_1 \leq M}$. Any IR, IC, and NPT mechanism $\mu = (q, s)$ satisfies $s \geq 0$ and $s(0) = 0$, and so $\mathbb{E}[s(X_{(M)})] = \mathbb{E}[s(X)1_{||X||_1 \leq M}]$ monotonically increases to $\mathbb{E}[s(X)]$ as $M$ increases to infinity. Therefore $\text{Rev}(X_{(M)})$ monotonically increases to $\text{Rev}(X)$. 30
If \( X^n \xrightarrow{D} X \) then \( X^n_{(M)} \xrightarrow{D} X_{(M)} \) for almost every \( M > 0 \)—specifically, for those \( M \) where \( \mathbb{P}[\|X\|_1 = M] = 0 \)—and so \( \lim_n \text{REV}(X^n_{(M)}) = \text{REV}(X_{(M)}) \) by Proposition 11. Now \( \text{REV}(X^n) \geq \text{REV}(X^n_{(M)}) \), and hence \( \lim\inf_n \text{REV}(X^n) \geq \text{REV}(X_{(M)}) \) for almost every \( M \). Letting \( M \to \infty \) proves the first part of the theorem.

If in addition the sequence \( X_n \) is uniformly integrable, then for every \( \varepsilon > 0 \) there is \( M > 0 \) with \( \mathbb{P}[\|X\|_1 = M] = 0 \) that is large enough so that \( \mathbb{E}[\|X^n\|_1 \mathbf{1}_{\|X^n\|_1 > M}] \leq \varepsilon \) for all \( n \). Since, as seen above, \( 0 \leq s(x) \leq \|x\|_1 \) for every IR and NPT mechanism \( \mu = (q, s) \), it follows that \( \mathbb{E}[s(X^n)\mathbf{1}_{\|X^n\|_1 > M}] \leq \varepsilon \) for all \( n \), and thus \( \text{REV}(X^n_{(M)}) \geq \text{REV}(X^n) - \varepsilon \) for all \( n \) (this also shows that the revenues are all finite, as they are bounded by \( M + \varepsilon \)). Therefore \( \text{REV}(X) \geq \text{REV}(X_{(M)}) \geq \lim_n \text{REV}(X^n_{(M)}) \geq \lim\sup_n \text{REV}(X^n) - \varepsilon \), which, together with the first part of the theorem, proves the second part.

### A.1 Continuous Valuations

It is often convenient—as in the present paper—to restrict attention to random valuations whose distributions admit a density function (i.e., their cumulative distribution functions are absolutely continuous; we refer to these for now as “continuous”). We now show that for the results in the present paper one may restrict attention to continuous random valuations. Indeed, assume that we have already proved a result of the form \( \text{REV}(X) \leq \theta \text{SREV}(X) \) for all such continuous \( X \), and let \( X \) be a valuation that is not necessarily continuous (and so it may have atoms and even finite support). First, because, as we have seen in the proof of Theorem 12 above, \( \text{REV}(X_{(M)}) \to_{M \to \infty} \text{REV}(X) \) and \( \text{SREV}(X_{(M)}) \to_{M \to \infty} \text{SREV}(X) \), it suffices to prove the result for random valuations \( X \) with bounded values, say \( \|X\|_1 \leq M \). Second, let \( U \) be independent of \( X \) and distributed uniformly on \([0, 1]^k\), and for every \( n \) define \( X^n := X + (1/n)U \). Then, clearly, the valuations \( X^n \) are continuous, \( X^n \xrightarrow{D} X \),

\[31\text{These are the points of continuity of the cumulative distribution function of } \|X\|_1; \text{ see Corollary 1 to Theorem 5.1 in Billingsley (1968).}\]
and the sequence \( X^n \) is bounded \((||X^n||, 1 \leq ||X||, 1 + (1/n)k \leq M + k)\); therefore \( \text{Rev}(X^n) \rightarrow_{n \rightarrow \infty} \text{Rev}(X) \) and \( \text{SRev}(X^n) \rightarrow_{n \rightarrow \infty} \text{SRev}(X) \) (apply the second part of Theorem 12 to the sequences \( X^n \overset{D}{\rightarrow} X \) and \( X^n_i \overset{D}{\rightarrow} X_i \) for all goods \( i \)). Thus \( \text{Rev}(X) \leq \theta \text{SRev}(X) \) holds for every bounded \( X \), and so for every \( X \).

\[ \text{B Appendix: Nonsymmetric Diagonals} \]

In this appendix we illustrate how the use of nonsymmetric diagonals alone may strictly improve the 50% bound of Hart and Nisan (2017), and, in some cases, also the 62% bound of our Theorem 7. However, this improvement is not uniform, in the sense that it does not yield a better constant than 50%.

**Proposition 13** Let \( X = (X_1, X_2) \) be a two-good random valuation with independent goods. Then

\[
\text{Rev}(X_1, X_2) \leq \left( \sqrt{\text{Rev}(X_1)} + \sqrt{\text{Rev}(X_2)} \right)^2.
\]

**Remark.** When \( \text{Rev}(X_1) \neq \text{Rev}(X_2) \) the right-hand side is strictly less than \( 2(\text{Rev}(X_1) + \text{Rev}(X_2)) \), the bound of Theorem A of Hart and Nisan (2017) (when \( \text{Rev}(X_1) = \text{Rev}(X_2) \) the two bounds are the same).

**Proof.** We follow the proof of Theorem A in Hart and Nisan (2017), but we now split the computation along the diagonal \( Y = \lambda Z \) for some \( \lambda > 0 \)

\[ \text{Rev}(X_1, X_2) \leq R_1 + R_2 + \min \left\{ \frac{2}{\sqrt{e}} \sqrt{R_1 R_2}, \frac{2}{e} \sqrt{R_1 R_2} + \left( 1 - \frac{1}{e} \right) \min \{R_1, R_2\} \right\}, \]

where \( R_i = \text{Rev}(X_i) \) for \( i = 1, 2 \).

\[32\]
(instead of splitting along $Y = Z$). The arguments in the proof there carry over, and, for each fixed value $z$ of $Z$, we now have

$$
\mathbb{E}[s(Y, Z)1_{Y \geq \lambda z} | Z = z] \leq \text{Rev}(Y) + z \mathbb{P}[Y \geq \lambda z] = \text{Rev}(Y) + \frac{1}{\lambda}(\lambda z) \mathbb{P}[Y \geq \lambda z] \\
\leq \text{Rev}(Y) + \frac{1}{\lambda}\text{Rev}(Y) = \left(1 + \frac{1}{\lambda}\right)\text{Rev}(Y).
$$

Similarly, for each fixed value $y$ of $Y$,

$$
\mathbb{E}[s(Y, Z)1_{Z \geq \lambda(1/\lambda)y} | Y = y] \leq \text{Rev}(Z) + y \mathbb{P}[Z \geq \frac{y}{\lambda}] = \text{Rev}(Z) + \frac{y}{\lambda}\mathbb{P}[Z \geq \frac{y}{\lambda}] \\
\leq \text{Rev}(Z) + \lambda\text{Rev}(Z) = (1 + \lambda)\text{Rev}(Z).
$$

Taking expectation over the values of $Z$ and $Y$, adding the two inequalities, and then minimizing the resulting expression over $\lambda$ (by taking $\lambda = \sqrt{\text{Rev}(Y)/\text{Rev}(Z)}$ ) yields the result.

**Remark.** A better bound than the one of Proposition 13, albeit also non-uniform, has been obtained by Kupfer (2017).

**References**

Babaioff, M., N. Immorlica, B. Lucier, and S. M. Weinberg (2014), “A Simple and Approximately Optimal Mechanism for an Additive Buyer,” *FOCS 2014: Proceedings of the 55th Annual Symposium on Foundations of Computer Science*, 21–30.

Babaioff, M., N. Nisan, and A. Rubinstein (2018), “Optimal Deterministic Mechanisms for an Additive Buyer,” *EC 2018: Proceedings of the 19th ACM Conference on Economics and Computation*, 429.

Billingsley, P. (1968), *Convergence of Probability Measures*, Wiley.
Briest, P., S. Chawla, R. Kleinberg, and M. Weinberg (2015), “Pricing Randomized Allocations,” *Journal of Economic Theory* 156, 144–174.

Chawla, S., D. L. Malec, and B. Sivan (2010), “The Power of Randomness in Bayesian Optimal Mechanism Design,” *EC 2010: Proceedings of the 11th ACM Conference on Electronic Commerce*, 149–158.

Daskalakis, C., A. Deckelbaum, and C. Tzamos (2017), “Strong Duality for a Multiple-Good Monopolist,” *Econometrica* 85, 735–767.

Hart, S. and N. Nisan (2013), “The Menu-Size Complexity of Auctions,” The Hebrew University of Jerusalem, The Hebrew University of Jerusalem, Center for Rationality DP-637; arXiv 1304.6116; *EC 2013: Proceedings of the 14th ACM Conference on Electronic Commerce*, 565–566.

Hart, S. and N. Nisan (2017), “Approximate Revenue Maximization with Multiple Items,” *Journal of Economic Theory* 172, 313–347; *EC 2012: Proceedings of the 13th ACM Conference on Electronic Commerce*, 656.

Hart, S. and P. J. Reny (2015), “Maximal Revenue with Multiple Goods: Nonmonotonicity and Other Observations,” *Theoretical Economics* 10, 893–922.

Kupfer, R. (2017), “A Note on Approximate Revenue Maximization with Two Items,” arXiv 1712.03518.

Monteiro, P. K. (2015), “A Note on the Continuity of the Optimal Auction,” *Economics Letters* 137, 127–130.

Myerson, R. B. (1981), “Optimal Auction Design,” *Mathematics of Operations Research* 6, 58–73.

Rockafellar, R. T. (1970), *Convex Analysis*, Princeton University Press.