RESEARCH ARTICLE

Efficacy and Safety of Non-Vitamin K Antagonist Oral Anticoagulants versus Vitamin K Antagonist Oral Anticoagulants in Patients Undergoing Radiofrequency Catheter Ablation of Atrial Fibrillation: A Meta-Analysis

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Abstract

Background
Use of the non-vitamin K antagonist oral anticoagulants (NOACs) is endorsed by current guidelines for stroke prevention in patients with atrial fibrillation (AF). However efficacy and safety of NOACs in patients undergoing catheter ablation (RFCA) of AF has not been well established yet.

Objectives
To perform a meta-analysis of all studies comparing NOACs and vitamin K antagonist oral anticoagulants (VKAs) in patients undergoing RFCA.

Data Sources
Studies were searched for in PubMed and Google Scholar databases.

Study Eligibility Criteria
Studies were considered eligible if: they evaluated the clinical impact of NOACs versus VKAs; they specifically analyzed the use of anticoagulants during periprocedural phase of RFCA; they reported clinical outcome data.
Study Appraisal and Synthesis Methods

25 studies were selected, including 9881 cases. The summary measure used was the risk ratio (RR) with 95% confidence interval (CI). The random-effects or the fixed effect model were used to synthesize results from the selected studies.

Results

There was no significant difference in thromboembolic complications (RR 1.39; p=0.13). Bleeding complications were significantly lower in the NOACs-treated arm as compared to VKAs (RR=0.67, p<0.001). Interestingly, a larger number of thromboembolic events was found in the VKAs-treated arm in those studies where VKAs had been interrupted during the periprocedural phase (RR=0.68; p = ns). In this same subgroup a significantly higher incidence of both minor (RR=0.54; p=0.002) and major bleeding (RR=0.41; p=0.01) events was recorded. Conversely, the incidence of thromboembolic events in the VKAs-treated arm was significantly lower in those studies with uninterrupted periprocedural anticoagulation treatment (RR=1.89; p=0.02).

Limitations

As with every meta-analysis, no patients-level data were available.

Conclusions and Implications

The use of NOACs in patients undergoing RFCA is safe, given the lower incidence of bleedings observed with NOACs. On the other side, periprocedural interruption of VKAs and bridging with heparin is associated with a higher bleeding rate with no significant benefit on onset of thromboembolic events.

Introduction

Atrial fibrillation (AF) is the most common arrhythmia, its prevalence in the developed World is approximately 1.5–2% of the general population [1]. AF is associated with a fivefold increase in the risk of thromboembolic stroke [2], hence prophylactic anticoagulation is a cornerstone in the clinical management of AF. Thromboembolic prophylaxis by means of oral anticoagulants was shown to be very effective, leading to a 60% relative risk reduction of stroke, compared to placebo [3]. However, vitamin K antagonists, the only available drugs for long-term anticoagulation over the last 60 years, have some flaws: a) several days are needed to reach the full therapeutic effect; b) complex overlap with parenteral anticoagulants; c) narrow therapeutic window, which potentially exposes to the dreadful risk of bleeding; d) significant interaction with several drugs and food, making continuous monitoring of the therapeutic effect mandatory; e) common genetic variations of its metabolism. For these reasons, up to 65% of all patients at risk were not taking oral anticoagulants (OACs), while the international normalized ratio (INR) was out of range in a further 19% of patients [4]. Over the last years, non-vitamin K antagonist oral anticoagulant drugs (NOACs) have been developed, including direct thrombin inhibitors (dabigatran) and factor Xa inhibitors (rivaroxaban, apixaban). Their therapeutic use for prevention of cardio-embolic complications was validated in recent large phase III trials, demonstrating their non-inferiority, and even superiority, in some cases, to warfarin [5,6,7].
Therefore, use of NOACs is currently recommended by guidelines, along with vitamin K antagonists, for stroke prevention in patients with non-valvular atrial fibrillation [8]. However, giving the relatively short experience with this new class of OACs in clinical practice, clinicians still face uncertainties on their efficacy and safety in patients undergoing catheter ablation of atrial fibrillation, due to the lack of clinical evidence in this specific population. We therefore aim to systematically evaluate all available evidence on the use of NOACs versus vitamin K antagonist oral anticoagulants (VKAs)in patients with atrial fibrillation allocated to a rhythm management therapy by radiofrequency catheter ablation (RFCA) using a meta-analytic approach to synthesize the results from all available studies. In particular, specific objectives of the present meta-analysis were: 1) evaluate the efficacy and safety of NOACs versus VKAs in patients with atrial fibrillation undergoing RFCA, and 2) test whether peri-procedural interruption of VKAs with heparin bridging could affect the clinical outcome of patients undergoing RFCA.

Methods

Search strategy and study selection

Published studies comparing NOACs to VKAs were searched for within PubMed and Google Scholar electronic databases up to February 28th 2015. The following key words were used: “dabigatran” “rivaroxaban” “apixaban” “atrial fibrillation”, and “radiofrequency catheter ablation”. Time of publication was not a limiting criterium for our analysis, while an English language restriction was made. All reports were independently screened by two investigators for eligibility. In addition, references from the selected articles were scanned for eligible studies, as already described before [9]. Any disagreement was resolved through discussion. All included studies were thoroughly checked and classified to exclude duplicity of data. Studies were considered eligible if: a) they evaluated the clinical impact of NOACs versus VKAs; b) they specifically analyzed the use of anticoagulants during periprocedural phase of RFCA; c) they reported clinical outcome data (stroke, transient ischemic attack or systemic embolism, major or minor bleedings). Exclusion criteria were (just one was sufficient for study exclusion): duplicate publication, pre-specified endpoint measure not specified.

Data abstraction and validity assessment

Baseline characteristics and numbers of events were extracted from all studies by two independent reviewers. Divergences were resolved by consensus after discussion. The following data were abstracted: year of publication, study location, number of patients, study design, clinical outcome data (stroke, systemic or pulmonary embolism, transient ischemic attack, minor bleedings and major bleedings), and baseline patients’ characteristics. In studies in which propensity score matching was applied to generate a cohort of two groups with balanced baseline data, only “after propensity score matching” data were considered in meta-analysis. Outcome data were separately analyzed for those studies reporting independent comparisons between multiple groups. Selection and data abstraction were performed according to the PRISMA statement (S1 Table) [10].

Statistical analysis

The summary measure used was the risk ratio (RR) with 95% confidence interval (CI). The random-effects model described by DerSimonian and Laird was used to combine the results from single studies [11]. This model calculates a weighted average by incorporating within-study and between-study variations. The Mantel-Haenszel method (fixed effect model) was
also used to assess the effect of model assumptions on our conclusions, depending on study heterogeneity [12]. Forest plots were used to graphically display the results of the meta-analysis, as already previously described [13]. Briefly, the left column lists the names of the studies. The right column is a plot of the measure of effect (risk ratio) for every study (represented by a square), incorporating 95% confidence intervals, represented by horizontal lines. The area of each square is proportional to the study’s weight in the meta-analysis. Finally, the central column shows the calculation results of the meta-analysis. The overall effect measure is reported on the bottom line of the plot (diamond), where the width indicates the confidence interval. Heterogeneity was assessed with the Cochrane Q test using a chi-squared function (p values < 0.10 were considered significant). In addition, I² values were calculated to estimate the proportion of variation in the effect size attributable to between-study heterogeneity. An I² value > 20% was considered significant. A random-effects model was preferred if appropriate, with a two-sided p value < 0.05 considered significant, as previously described [14]. Small study effects were evaluated through graphical inspection of funnel plots and statistically using the Egger’s test [15]. Analysis was performed by means of Excel spreadsheets and RevMan 5.

Results
Study characteristics

Our database search retrieved a total of 77 entries from PubMed, 10 from Google Scholar and 5 from congress abstracts. Overall, a total of 77 studies were available after removal of duplicates, which were screened with exclusion of further 52 studies based on pre-specified criteria. In particular, most studies were excluded at this stage while they either did not report the clinical endpoints, they were review articles, they were written in non-English language. Finally, a total of 25 studies fulfilled the inclusion criteria and were selected for this meta-analysis. The step-wise selection process is illustrated in Fig 1. Details of the selected studies are showed in table 1. Thromboembolic and bleeding risk profiles were similar in both study arms for all studies, except three [16,17,18]. Data on the CHADS2 or the CHA2DS2-VASc score and the HAS-BLED score are showed in S2 Table. Endpoints’ definition for all selected studies are reported in S3 Table.

Periprocedural anticoagulation was managed in different ways in the various studies. In fact, VKAs administration was interrupted in some studies and kept uninterrupted in other studies. In the first case, the anticoagulation treatment was stopped 2 to 5 days before the RFCA procedure and replaced by low molecular weight heparin (LMWH) bridging. In the remaining studies VKAs administration remained uninterrupted throughout the procedural phase, keeping the INR value within the 2–3 target. On the contrary, periprocedural NOAC management was homogeneous across almost all studies included. The NOAC assumption was stopped 12 to 30 h before the RFCA procedure without any LMWH bridging and started again 6 to 24 h after. Finally, in both study arms (VKAs and NOACs) intraprocedural anticoagulation was achieved with the administration of a weight-based unfractionated heparin bolus just after crossing the interatrial septum to the left atrium with the transseptal sheath. After bolus administration, iv heparin infusion (1000 IU/h) was continued during the procedure, keeping the activating clotting time (ACT) between 250 and 400 seconds. Of all selected studies, one is a post-hoc analysis of a large randomized trial [19]. It evaluated efficacy and safety of rivaroxaban in AF patients treated with either cardioversion or RFCA. Sixteen studies [16,17,20,21,22,23,24,25,26,27,28,29,30,31,32,33] are retrospective analyses including patients that had undergone RFCA of AF with periprocedural anticoagulation utilizing warfarin and dabigatran or rivaroxaban. Two studies [34,35] are retrospective analyses including patients that had undergone left atrial RFCA with periprocedural anticoagulation utilizing...
phenprocoumon and rivaroxaban or apixaban. One study is a small-randomized prospective trial that evaluated the safety and efficacy of dabigatran in comparison with warfarin in the setting of RFCA of AF [36]. The remaining five studies are prospective, non-randomized, observational studies analyzing the safety and efficacy of dabigatran/rivaroxaban in comparison with warfarin in patients undergoing RFCA of AF [18,37,38,39,40]. Study quality assessment performed according to the PRISMA statement.
Table 1. Details of the selected studies included in the Meta-analysis.

| STUDY     | YEAR | LOCATION   | N   | SOURCE TYPE    | STUDY DESIGN                      | PRIMARY END POINT                                          | OTHER END POINTS                                    | FUP       |
|-----------|------|------------|-----|----------------|-----------------------------------|-----------------------------------------------------------|------------------------------------------------------|-----------|
| Kaseno [20] | 2012 | Japan      | 211 | Journal Article | Retrospective non-blinded single centre study | Efficacy and safety of peri-procedural dabigatran in AF ablation. | Death, Stroke, Major and Minor bleeding | 30 days   |
| Snipelisky [21] | 2012 | USA        | 156 | Journal Article | Retrospective single centre study | Provide information on the use of dabigatran in patients undergoing atrial arrhythmias ablation | Major and minor Bleeding | 30 days   |
| Lakkireddy [37] | 2012 | USA        | 290 | Journal Article | Prospective study | Feasibility and safety of continuing dabigatran during AF ablation. | Major and minor bleeding | 30 days   |
| Konduru [22] | 2012 | USA        | 76  | Journal Article | Retrospective study | Compare heparin administration and degree of anticoagulation achieved in patients treated with dabigatran and warfarin before AF ablation | Stroke, Major Bleeding | 30 days   |
| Mendoza [28] | 2012 | USA        | 119 | Conference Abstract | Retrospective study | Feasibility and safety of dabigatran in the setting of AF ablation. | Stroke, major and minor bleeding | 30 days   |
| Rowley [29] | 2012 | USA        | 282 | Conference Abstract | Retrospective study | Feasibility and safety of dabigatran for AF ablation. | Stroke, major and minor bleeding | 30 days   |
| Ellis [31] | 2012 | USA        | 171 | Conference Abstract | Retrospective study | Feasibility and safety of dabigatran for AF ablation. | Stroke, major and minor bleeding | 30 days   |
| Pavaci [30] | 2012 | Germany    | 54  | Conference Abstract | Retrospective study | Feasibility and safety of dabigatran for AF ablation. | Stroke, major and minor bleeding | 30 days   |
| Kaiser [23] | 2013 | USA        | 257 | Journal Article | Retrospective study | Feasibility and safety of dabigatran for AF ablation. | Stroke and Major bleeding | 90 days   |
| Haines [24] | 2013 | USA, Canada | 404 | Journal Article | Retrospective multicentre pilot trial | Feasibility and safety of dabigatran for AF ablation. | Freedom from stroke and major bleeding | -         |
| Imamura [38] | 2013 | Japan      | 227 | Journal Article | Prospective observational single centre study | Feasibility and safety of dabigatran for AF ablation. | Stroke, major and minor bleeding | 30 days   |
| Maddox [16] | 2013 | USA        | 463 | Journal Article | Retrospective single centre study | Feasibility and safety of dabigatran for AF ablation. | Stroke, major and minor bleeding | 90 days   |
| Nin [36] | 2013 | Japan      | 90  | Journal Article | Randomized single centre study | Feasibility and safety of dabigatran for AF ablation. | Stroke, major and minor bleeding | 12 months |
| Bernard [26] | 2013 | USA        | 274 | Conference Abstract | Retrospective single centre study | Compare the feasibility and safety of dabigatran, rivaroxaban and warfarin, used for peri-procedural anticoagulation in AF ablation | Stroke, major and minor bleeding | 30 days   |
| Kim [26] | 2013 | USA        | 763 | Journal Article | Retrospective, not randomized case control study | Compare the feasibility and safety of dabigatran and warfarin used for peri-procedural anticoagulation in AF ablation | Thromboembolism, major and minor bleeding | 90 days   |

(Continued)
Clinical events

The 25 studies selected for this meta-analysis included a total of 9881 patients (4229 treated with NOACs and 5652 treated with VKAs). A total of 79 embolic events (35 in the NOACs...
arm and 44 in the VKAs arm) and a total of 533 bleeding events (182 in the NOACs arm and 351 in the VKAs arm) were registered.

### Stroke and thromboembolism

Analyzing the results from all studies reporting on thromboembolic events, no significant difference in the incidence of this endpoint was found (RR 1.39, 95%CI 0.91–2.14; p = 0.13) (Fig 2A). Of note, the sensitivity analysis revealed that this result was constant across all selected studies and the overall effect didn’t significantly change even after removal of single studies, with one exception. In fact, we evaluated the influence of removal of each of the two studies where the two treatment arms were slightly unbalanced in the baseline predicted thromboembolic risk. Indeed, removal of the study by Kaiser and colleagues [23] had no influence on the cumulative result while removal of the study by Providencia and colleagues [40] significantly affected the cumulative result, with a significantly lower incidence of the endpoint in the VKAs-treated group (RR 1.61, 95%CI 1.02–2.52; p = 0.04). To further investigate the origin of the observed incidence of the endpoint, we performed another subgroup analysis, dividing the studies in two groups: the first (interrupted vitamin K antagonist oral anticoagulants, IVKAs) including all studies in which periprocedural VKAs interruption with heparin bridging was performed; the second (uninterrupted vitamin K antagonist oral anticoagulants, CVKAs) including the studies in which no periprocedural interruption of VKAs treatment was applied.

Of note, this subgroup analysis showed a significant difference in the incidence of thromboembolic events in favor of VKAs, that was limited to the subgroup of studies in which the VKAs treatment had not been interrupted during the periprocedural phase (RR = 1.89 95% CI 1.13–3.18; p = 0.02), while no significant difference was found in the interrupted group (RR = 0.68 95% CI 0.32–1.47; p = 0.33) (Fig 2B). Furthermore, since 22 studies had used dabigatran and 6 rivaroxaban, we performed a further subgroup analysis, which did not show a substantial difference between the study subgroups using different NOACs (S1 Fig). Accordingly, comparison between the study subgroups didn’t show a statistically significant difference (p = 0.30). In addition to the composite endpoint, enough data were available from the selected studies, to look after the overall affect on single clinical endpoints. Interestingly, a similar result was obtained analyzing the summary effect on the cumulative endpoint of transient ischemic attack (TIA) or silent cerebral ischemia (SCI), where a significant difference was found in favor of the VKAs-treated group (RR = 2.45 95% CI 1.40–4.28; p = 0.002) (Fig 3A). This result was driven by the subgroup of studies in which VKAs were not interrupted during the periprocedural phase (RR = 2.42 95% CI 1.30–4.49; p = 0.005), while the difference in favor of VKAs lost its statistical significance when the analysis was restricted to the studies in which the oral anticoagulation with VKAs had been interrupted and bridged with heparin during the periprocedural phase (RR = 2.57 95% CI 0.70–9.47; p = 0.16) (Fig 3A). A smaller number of studies provided data on the isolated endpoint of ischemic stroke [23,24,28,29,30,40]. Summing the evidence from these studies, no difference was found between NOACs and VKAs-treated arms (RR 0.77, 95% CI 0.29–2.01, p = 0.59) (Fig 3B).

### Bleedings

Analyzing the results from all studies reporting this endpoint, a lower incidence of total bleeding events was registered in the NOACs-treated arm (RR = 0.67, 95% CI 0.54–0.84; p = 0.0004) (Fig 4A). A subgroup analysis showed a substantial difference in the incidence of total bleedings in favor of the NOACs-treated arm both in the subgroup of studies in which the VKAs treatment had been interrupted and a heparin bridging had been performed during the periprocedural phase (RR = 0.53 95% CI 0.35–0.80; p = 0.002) (Fig 4B) and in the subgroup with
Fig 2. Meta-analysis of difference in thromboembolic events. Panel A: forest plot and summary effect of the difference in the incidence of thromboembolic events. Panel B: difference in the thromboembolic events is shown in the subgroups of patients undergoing interrupted vs. uninterrupted periprocedural VKAs.

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### TIA or Silent ischemia

| Study or Subgroup | NOACs Events | VKAs Events | Total | Weight | Risk Ratio IV, Random, 95% CI | Risk Ratio IV, Random, 95% CI |
|-------------------|--------------|-------------|-------|--------|-----------------------------|-----------------------------|
|                   |              |             |       |        |                             |                             |
| 3.1.1 VKAs Uninterrupted |              |             |       |        |                             |                             |
| Diller            | 272          | 272         |       |        |                             |                             |
| Haines            | 202          | 202         |       |        |                             |                             |
| Ichiki            | 30           | 180         | 57.5% |        |                             |                             |
| Kaess             | 105          | 210         |       |        |                             |                             |
| Kaiser            | 122          | 135         |       | 3.1%   |                             |                             |
| Kaseno            | 110          | 101         |       | 4.1%   |                             |                             |
| Lakireddy         | 145          | 145         |       | 3.6%   |                             |                             |
| Maddox            | 212          | 251         |       | 3.1%   |                             |                             |
| Stepanyan – Dabi  | 89           | 124         |       | 4.1%   |                             |                             |
| Stepanyan – Riva  | 98           | 114         |       | 3.1%   |                             |                             |
| **Subtotal (95% CI)** | 1385        | 1724        | 81.6% |        | 2.42 [1.30, 4.49]           |                             |
| Total events      | 16           | 21          |       |        |                             |                             |
| Heterogeneity: Tau^2 = 0.00; Chi^2 = 2.62, df = 7 (P = 0.92); I^2 = 0% |                             |                             |
| Test for overall effect: Z = 2.80 (P = 0.005) |                             |                             |

| 3.1.2 VKAs Interrupted |              |             |       |        |                             |                             |
| Imamura            | 101          | 126         |       | 3.1%   |                             |                             |
| Providencia – Dabi | 117          | 192         |       | 3.1%   |                             |                             |
| Providencia – Riva | 188          | 192         |       | 3.1%   |                             |                             |
| Rowley             | 113          | 169         |       | 3.1%   |                             |                             |
| Winkle – Dabi      | 426          | 1113        |       | 3.1%   |                             |                             |
| Winkle – Riva      | 187          | 1113        |       | 3.1%   |                             |                             |
| **Subtotal (95% CI)** | 1191        | 2905        | 18.4% |        | 2.57 [0.70, 4.47]           |                             |
| Total events       | 4            | 2           |       |        |                             |                             |
| Heterogeneity: Tau^2 = 0.00; Chi^2 = 0.67, df = 5 (P = 0.98); I^2 = 0% |                             |                             |
| Test for overall effect: Z = 1.42 (P = 0.16) |                             |                             |

### Stroke

| Study or Subgroup | NOACs Events | VKAs Events | Total | Weight | Risk Ratio IV, Random, 95% CI | Risk Ratio IV, Random, 95% CI |
|-------------------|--------------|-------------|-------|--------|-----------------------------|-----------------------------|
|                   |              |             |       |        |                             |                             |
| Diller            | 272          | 272         |       |        |                             |                             |
| Haines            | 202          | 202         |       |        |                             |                             |
| Kaess             | 105          | 210         |       |        |                             |                             |
| Kaiser            | 122          | 135         |       | 9.1%   |                             |                             |
| Pavaci            | 27           | 27          |       | 9.3%   |                             |                             |
| Providencia – Dabi| 176          | 192         |       | 10.9%  |                             |                             |
| Providencia – Riva| 188          | 192         |       | 19.4%  |                             |                             |
| Rowley            | 113          | 169         |       | 12.1%  |                             |                             |
| Winkle – Dabi     | 426          | 1113        |       | 19.3%  |                             |                             |
| Winkle – Riva     | 187          | 1113        |       | 10.9%  |                             |                             |
| **Total (95% CI)** | 1818         | 3625        | 100.0%|        | 0.77 [0.29, 2.01]           |                             |
| Total events      | 6            | 17          |       |        |                             |                             |
| Heterogeneity: Tau^2 = 0.00; Chi^2 = 5.00, df = 7 (P = 0.66); I^2 = 0% |                             |                             |
| Test for overall effect: Z = 0.54 (P = 0.59) |                             |                             |

**Fig 3. Meta-analysis of difference in TIA and Stroke.** Panel A: forest plot and summary effect of the difference in the incidence of TIA or silent ischemia. Panel B: forest plot and summary effect of the difference in the incidence of Stroke.

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no periprocedural VKAs interruption, where a much narrower but still statistically significant difference was observed (RR = 0.78 95% CI 0.61–1.00; p = 0.05) (Fig 4B). Furthermore, since 22 studies had used dabigatran and 6 rivaroxaban, we performed a further subgroup analysis, which did not show a substantial difference between the study subgroups using different
**Fig 4. Meta-analysis of difference in total bleeding events.** Panel A: forest plot and summary effect of the difference in the incidence of total bleedings. Panel B: difference in the total bleeding events is shown in the subgroups of patients undergoing interrupted vs. uninterrupted periprocedural VKAs.

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NOACs (S1 Fig). Accordingly, comparison between the study subgroups didn’t show a statistically significant difference (p = 0.66).

However, when the same comparison was performed after separating major (intracranial, gastrointestinal or any other type of bleedings requiring blood transfusion) from minor bleedings, we found a significantly lower incidence in minor bleedings in favor of the NOACs (RR = 0.69 95% CI 0.56–0.85; p = 0.0003) (Fig 5A), while no statistically significant difference in the incidence of major bleedings was evident, despite a clear trend in favor of the NOACs-treated arm (RR = 0.82 95% CI 0.53–1.25; p = 0.36) (Fig 6). Interestingly, interruption of VKAs during the periprocedural phase (with concomitant heparin bridging) was associated to a higher incidence of both minor (RR = 0.54 95% CI 0.36–0.80; p = 0.002) (Fig 5B) and major bleedings (RR = 0.41 95% CI 0.20–0.82; p = 0.01) (Fig 6) in comparison to the treatment with NOACs.

**Discussion**

The present study is the largest and most comprehensive meta-analysis comparing the non-vitamin K antagonists oral anticoagulants to vitamin K antagonist oral anticoagulants in patients undergoing rhythm control management of atrial fibrillation by RFCA, with 25 selected studies that included 9881 cases. In contrast with a recently published meta-analysis, performed on a smaller group of studies [41], we found no significant difference in the incidence of total thromboembolic events between the NOACs-treated arm and the VKAs group. Foremost, the present meta-analysis is the first showing an advantage in favor of the VKAs-treated group on thromboembolic events in the subgroup of studies in which the VKAs treatment wasn’t interrupted in the periprocedural phase. Moreover, in the same study subgroup, interruption of VKAs with concomitant heparin bridging apparently exposed the patients to a higher bleeding risk. These latter results are not surprising, since a recent randomized international multicentre trial demonstrated a better outcome when the RFCA procedure was performed without VKAs discontinuation, over the alternative strategy of discontinuing VKAs with concomitant heparin bridging [42]. Previous meta-analyses performed on smaller number of studies, have shown a similar efficacy and safety profile of NOACs and VKAs [43,44,45]. In particular, Providencia et al. [44] performed a meta-analysis and reported no significant differences between dabigatran and warfarin for thromboembolic events (0.55% with dabigatran vs. 0.17% with warfarin, risk ratio 1.78). Similarly, Hohnloser et al. [45] reported no significant differences. With both these meta-analyses a numerically lower recurrence of thromboembolic events was observed within the warfarin-treated arm. Compared to previous meta-analyses, however, our study includes the largest number of studies and patients available so far assigned to a rhythm management strategy of atrial fibrillation, by means of RFCA, including dabigatran and rivaroxaban, while all previous analyses were mostly focused on dabigatran-treated patients. The novel finding of an increased incidence of thromboembolic events in patients undergoing RFCA in the NOACs-treatment arm could be related to the different pharmacodynamic profile of these two classes of oral anticoagulants. In fact, while the anticoagulant effect of VKAs gradually fluctuates over longer periods, warranting a seemingly stable anticoagulant effect on the short term in patients on target, the half-life of NOACs ranges 12–17h. Hence, their anticoagulant effect mostly vanishes within 24–36 hours from the last assumption. Therefore, the differences in the timing of administration of NOACs before and after RFCA, could have affected their efficacy in the periprocedural setting. In fact, in almost all studies, the NOAC was interrupted at least 12h before the RFCA procedure and started again the day after, allowing a short—yet not clinically irrelevant—window of suboptimal anticoagulation. This hypothesis is further supported by our findings in the subgroup analysis—which we could perform given the large number of studies.
Fig 5. Meta-analysis of difference in minor bleeding events. Panel A: forest plot and summary effect of the difference in the incidence of minor bleedings. Panel B: difference in the minor bleeding events is shown in the subgroups of patients undergoing interrupted vs uninterrupted periprocedural VKAs.

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included—showing that the superiority of VKAs in the prevention of thromboembolic events was less evident in the subgroup of studies in which the VKAs treatment was interrupted during the periprocedural phase. Altogether these findings highlight the importance of an optimal periprocedural coagulation, since thromboembolic events are concentrated around the time of rhythm conversion. At the same time, the higher bleeding rate observed with VKAs in the study subgroup with VKAs interruption and heparin bridging suggests that maintenance of a constant level of anticoagulation is of key importance in preventing dreadful bleeding events. Another possible explanation for the superiority observed with VKAs in the prevention of thromboembolic events is the chance to monitor their anticoagulant effect, while no validated method for assessing the anticoagulant efficacy of an ongoing treatment with NOACs is

### Major Bleedings

| Study or Subgroup | NOACs Events | VKAs Events | Risk Ratio | Risk Ratio |
|-------------------|--------------|-------------|------------|------------|
|                   | Total        | Total       | IV, Random, 95% CI | IV, Random, 95% CI |
| **5.1.1 VKAs Uninterrupted** |
| Arshad Uninterrupted | 2 374 | 10 276 | 5.1% | 0.15 [0.03, 0.67] |
| Basslouny | 4 344 | 5 344 | 6.1% | 0.80 [0.22, 2.95] |
| Bernard – Dabi | 2 155 | 2 44 | 3.6% | 0.28 [0.04, 1.96] |
| Bernard – Riva | 1 75 | 2 44 | 2.6% | 0.29 [0.03, 3.14] |
| Dillier | 1 272 | 1 272 | 2.0% | 1.00 [0.06, 15.91] |
| Ellis | 1 61 | 5 110 | 3.2% | 0.36 [0.04, 3.02] |
| Haines | 2 202 | 2 202 | 3.6% | 1.00 [0.14, 7.03] |
| Ichiki | 4 30 | 5 180 | 6.4% | 4.80 [1.37, 16.87] |
| Kaess | 1 105 | 1 210 | 2.1% | 2.00 [0.13, 31.66] |
| Kaiser | 2 122 | 1 135 | 2.6% | 2.21 [0.20, 24.10] |
| Kasedo | 0 110 | 2 101 | 1.8% | 0.18 [0.01, 3.78] |
| Kim | 4 191 | 12 572 | 7.2% | 1.00 [0.33, 3.06] |
| Konduru | 1 24 | 0 52 | 1.6% | 6.36 [0.27, 150.67] |
| Lakireddy | 9 145 | 1 145 | 3.3% | 9.00 [1.16, 70.13] |
| Maddox | 1 212 | 3 251 | 2.9% | 0.39 [0.04, 3.77] |
| Mendoza | 1 60 | 0 58 | 1.6% | 2.90 [0.12, 69.81] |
| Piccin | 30 160 | 21 161 | 11.8% | 1.44 [0.86, 2.40] |
| Stepairy – Dabi | 4 89 | 2 114 | 4.5% | 2.56 [0.48, 13.67] |
| Stepairy – Riva | 0 98 | 2 114 | 1.8% | 0.23 [0.01, 4.78] |
| Yamaji Uninterrupted | 0 106 | 2 203 | 1.8% | 0.38 [0.02, 7.87] |
| **Subtotal (95% CI)** | 2935 | 3588 | 75.6% | 1.03 [0.63, 1.68] |

**Total events** | 70 | 79 | Heterogeneity: $\tau^2 = 0.33; \chi^2 = 28.10$, df = 19 ($P = 0.08$); $I^2 = 32\%$

Test for overall effect: $Z = 0.10$ ($P = 0.92$)

### 5.1.2 VKAs Interrupted

| Arshad Interrupted | 2 374 | 5 232 | 4.6% | 0.25 [0.05, 1.27] |
| Imamura | 3 101 | 4 126 | 5.3% | 0.94 [0.21, 4.09] |
| Providencia – Dabi | 2 176 | 8 192 | 5.0% | 0.27 [0.06, 1.27] |
| Providencia – Riva | 3 188 | 8 192 | 6.1% | 0.38 [0.10, 1.42] |
| Rowley | 0 113 | 1 169 | 1.6% | 0.50 [0.02, 12.09] |
| Yamaji Interrupted | 0 106 | 2 194 | 1.8% | 0.36 [0.02, 7.52] |
| **Subtotal (95% CI)** | 1058 | 1105 | 24.4% | 0.41 [0.20, 0.82] |

**Total events** | 10 | 28 | Heterogeneity: $\tau^2 = 0.00; \chi^2 = 1.87$, df = 5 ($P = 0.87$); $I^2 = 0\%$

Test for overall effect: $Z = 2.52$ ($P = 0.01$)

### Total (95% CI)

| 3993 | 4693 | 100.0% | 0.82 [0.53, 1.25] |

**Total events** | 80 | 107 | Heterogeneity: $\tau^2 = 0.33; \chi^2 = 37.05$, df = 25 ($P = 0.06$); $I^2 = 33\%$

Test for overall effect: $Z = 0.92$ ($P = 0.36$)

Test for subgroup differences: $\chi^2 = 4.49$, df = 1 ($P = 0.03$), $I^2 = 77.7\%$

Fig 6. Meta-analysis of difference in major bleeding events. Forest plot and summary effect of the difference in the incidence of major bleedings within the subgroups of patients undergoing interrupted vs uninterrupted periprocedural VKAs. The bottom line reports the summary effect for both groups.

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available yet. This fact could also potentially explain the higher bleeding rate when VKAs treatment was periprocedurally substituted by LMWH bridging, which cannot be monitored. Finally, another aspect to take into account is represented by the possible drug interactions of dabigatran, the NOAC used in many studies, with amiodarone and dronedarone. Unfortunately, the frequency of use of these both drugs in the selected studies is not known. However, since they are frequently used in patients undergoing a rhythm management strategy, we cannot exclude that they were largely used in the selected studies. The results of the present meta-analysis underline the importance of an optimal balance between risks and benefits in medicine. In particular, during the periprocedural phase a higher (bleeding-) risk propensity is justified by the substantial prevention of thromboembolic events. However, these findings shouldn’t be generalized to other clinical settings. In fact, during the post-ablation phase, a lower thromboembolic risk probably doesn’t justify the routinely use of oral anticoagulation. A recent analysis from a large Danish cohort apparently supports this concept. In fact, use of oral anticoagulation beyond the third month after RFCA was associated with a 1.8% increase of serious bleedings, with only 0.6% reduction of the thromboembolic risk [46].

**Limitations**

Given the large number of non-randomized studies included in the present meta-analysis, we cannot exclude the risk of a selection bias. In fact only a minority of trials were randomized or prospective and a blinded analysis and central adjudication of endpoints was performed in a minority of the included studies. Nevertheless, the inclusion of non-randomized studies in a meta-analysis is recommended by many authors since it helps in generalizing the conclusions to the daily clinical routine. Furthermore, since the NOAC was interrupted 12–30h before the RFCA procedure and started again the day after in almost all studies, no comparison could be performed distinguishing interrupted and non-interrupted NOAC treatment. Finally, despite the studies were generally well balanced between the treatment arms, especially regarding the predicted thromboembolic risk and the bleeding risk, these last data were heterogeneously reported in the selected studies, making impossible to exploit them as a moderator variable for a meta-regression.

**Conclusions**

Notwithstanding the above cited potential limitations, the present meta-analysis—the largest and most comprehensive available to date—shows no significant difference for the prevention of thromboembolic complications in patients with atrial fibrillation managed with RFCA, although a trend is evident in favor of the VKAs-treated group. Interestingly, a significantly lower incidence of thromboembolic complications was found in favor of the VKAs-treated arm in the subgroup of studies in which the VKAs treatment was not interrupted in the periprocedural phase, although this was paid back with a slightly higher incidence of bleeding complications. Besides confirming the safety profile of NOACs, our results also highlight the importance of the optimal periprocedural management of anticoagulation. In this respect, the design of a prospective trial to evaluate the optimal periprocedural anticoagulation is desirable, to test efficacy and safety of uninterrupted periprocedural VKA versus uninterrupted periprocedural NOAC treatment in patients with atrial fibrillation undergoing a RFCA procedure.

**Supporting Information**

**S1 Fig. Dabigatran and Rivaroxaban Subgroups.** Subgroup analysis comparing the study subgroup using dabigatran to the study subgroup using rivaroxaban. (EPS)
S1 Table. Prisma 2009 checklist.

S2 Table. Baseline risk and events’ list. CHADS2 or the CHA2DS2-VASc score and the HAS-BLED score, along with the number of events for each single study.

S3 Table. Endpoints’ definition. Definition of thromboembolic and bleeding complications.

Author Contributions
Conceived and designed the experiments: GS SDR CI. Analyzed the data: SDR AP GS MM SG CI. Wrote the paper: GS SDR AP AC CI. Performed the analyses: AP SDR GS AC.

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