Vaccine schedules and the effect on humoral and intestinal immunity against poliovirus: a systematic review and network meta-analysis

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Summary

Background The eradication of wild and vaccine-derived poliovirus requires the global withdrawal of oral poliovirus vaccines (OPVs) and replacement with inactivated poliovirus vaccines (IPVs). The first phase of this effort was the withdrawal of the serotype 2 vaccine in April 2016, with a switch from trivalent OPVs to bivalent OPVs. The aim of our study was to produce comparative estimates of humoral and intestinal mucosal immunity associated with different routine immunisation schedules.

Methods We did a random-effect meta-analysis with single proportions and a network meta-analysis in a Bayesian framework to synthesise direct and indirect data. We searched MEDLINE and the Cochrane Library Central Register of Controlled Trials for randomised controlled trials published from Jan 1, 1980, to Nov 1, 2018, comparing poliovirus immunisation schedules in a primary series. Only trials done outside western Europe or North America and without variation in age schedules (ie, age at administration of the vaccine) between study groups were included in the analyses, because trials in high-income settings differ in vaccine immunogenicity and schedules from other settings and to ensure consistency within the network of trials. Data were extracted directly from the published reports. We assessed seroconversion against poliovirus serotypes 1, 2, and 3, and intestinal immunity against serotype 2, measured by absence of shedding poliovirus after a challenge OPV dose.

Findings We identified 437 unique studies; of them, 17 studies with a maximum of 8279 evaluable infants were eligible for assessment of humoral immunity, and eight studies with 4254 infants were eligible for intestinal immunity. There was low between-trial heterogeneity in the data (τ=0·05, 95% credible interval [CrI] 0·009–0·15). For serotype 2, the risk ratio (RR) of seroconversion after three doses of bivalent OPVs was 0·14 (95% CrI 0·11–0·17) compared with three doses of trivalent OPVs. The addition of one or two full doses of an IPV after a bivalent OPV schedule increased the RR to 0·85 (0·75–1·0) and 1·1 (0·98–1·4). However, the addition of an IPV to bivalent OPV schedules did not significantly increase intestinal immunity (0·33, 0·18–0·61), compared with trivalent OPVs alone. For serotypes 1 and 3, there was substantial inconsistency and between-trial heterogeneity between direct and indirect effects, so we only present pooled estimates on seroconversion, which were at least 80% for serotype 1 and at least 88% for serotype 3 for all vaccine schedules.

Interpretation For WHO’s polio eradication programme, the addition of one IPV dose for all birth cohorts should be prioritised to protect against paralysis caused by type 2 poliovirus; however, this inclusion will not prevent transmission or circulation in areas with faecal–oral transmission.

Funding UK Medical Research Council.

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Introduction In 1988, the World Health Assembly passed a resolution that committed WHO to eradicating poliomyelitis globally. The eradication effort has been centred on mass vaccination campaigns and achieving high routine immunisation coverage with Sabin oral poliovirus vaccines (OPVs) and Salk inactivated poliovirus vaccines (IPVs). More than 150 countries have relied on OPVs to eliminate poliovirus transmission and maintain a polio-free status; however, the cessation of all OPV use and replacement by IPVs is necessary because of the risk of vaccine-derived poliovirus and vaccine-associated paralytic poliomyelitis associated with the OPV. The first phase of this cessation was completed in April 2016, with the global withdrawal of the type 2 strain OPV and a switch from the trivalent OPV (tOPV) to a bivalent OPV (bOPV) formulation. Additionally, the Strategic Advisory Group of Experts on Immunisation (SAGE) recommended that at least one dose of IPV was introduced into all routine immunisation schedules to protect against poliomyelitis caused by serotype 2. However, constraints on IPV supply resulted in 39 countries having to delay IPV introduction or interrupt its routine use, with some countries adopting the use of intradermal fractional-dose IPV (fIPV).
Although network meta-analyses are used increasingly to compare interventions that have not been explicitly trialled. By contrast, network meta-analysis uses the quantitative relatedness (ie, relative effects) of interventions to estimate both the direct and indirect effects.9,10

We assessed the risk of bias with the Cochrane Collaboration’s tool for assessing bias in randomised trials and found it to be low-to-moderate for individual studies. The effect of various vaccine schedules on humoral and mucosal immunity has been addressed by many reviews and phase 3 clinical trials in different settings globally, creating a vast pool of recommendations for appropriate vaccine schedules.

Added value of this study
Schedules of bivalent OPV plus IPV have the potential to provide humoral immunity against serotype 2 poliovirus, which protects against paralysis from circulating vaccine-derived poliovirus of serotype 2. However, such a schedule will not induce intestinal protection against serotype 2, enabling transmission of this virus in some populations. There is little difference between the immunogenicity achieved through affordable IPV options (Salk IPV, Sabin IPV, fractional IPV, adjuvanted IPV, and monovalent IPV).

Methods
Search strategy and selection criteria
We did a systematic review and network meta-analysis of randomised controlled trials comparing the immunogenicity of different OPV and IPV routine immunisation schedules considered by WHO and member states in inducing humoral and intestinal immunity against poliovirus. This knowledge would be useful to inform global immunisation policy.

After termination of OPV use, the post-eradication schedule is planned to comprise a minimum of two IPV doses given after 14 weeks of age.4 There are several approaches to developing affordable IPV options, including restricting the number of IPV doses in routine immunisation to two, reducing the volume of each dose through intra­dermal administration, reducing the antigen content of each dose through use of adjuvants, and reducing the cost of production through developing IPV formulations to attenuated vaccines of the Sabin strains of poliovirus.7,8 These approaches have resulted in the development of alternative formulations to conventional intramuscular Salk IPV, including intradermal fIPV, adjuvanted IPV, monovalent type 2 IPV (mIPV2), and Sabin IPV (sIPV).

Accordingly, many clinical trials have evaluated the immunogenicity of different vaccine schedules. It is essential to develop a comprehensive overview of immunogenicity induced by different routine immunisation schedules against the three poliovirus serotypes. Standard meta-analysis approaches combine information from multiple studies to estimate the overall effectiveness of an intervention, but do not compare effectiveness between interventions that have not been explicitly trialled. By contrast, network meta-analysis uses the quantitative relatedness (ie, relative effects) of interventions to estimate both the direct and indirect effects.3-10 Although network meta-analyses are used increasingly to compare drugs, they have not been widely adopted to compare vaccine schedules.3,12

In this Article, we aim to estimate the relative immunogenicity of the different OPV and IPV routine immunisation schedules considered by WHO and member states in inducing humoral and intestinal immunity against poliovirus. This knowledge would be useful to inform global immunisation policy.

Evidence before this study
The phased removal of the oral polio vaccine (OPV) is occurring alongside introduction of the inactivated poliovirus vaccine (IPV), eradication of wild poliovirus cases, and prevention of the emergence and circulation of vaccine-derived polioviruses. We searched MEDLINE and the Cochrane Library Central Register of Controlled Trials for randomised controlled trials published from Jan 1, 1980, to Nov 1, 2018, that compare poliovirus immunisation schedules in primary series to compile in a network meta-analysis. We used the search terms: (polio OR poliovirus) AND vaccine AND (primary series OR routine OR infants) AND (seropositive OR seroconversion OR antibody OR mucosal immunity OR intestinal immunity). Only trials done outside western Europe or North America and without variation in age schedules between study groups were included in the analyses.

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vaccine AND (primary series OR routine OR infants) AND (seropositive OR seroconversion OR antibody OR mucosal immunity OR intestinal immunity). Trials were excluded if they were done in western Europe or North America, because of differences in vaccine immunogenicity and schedules used in these high-income settings, or if there was variation in age schedules (ie, age at administration of the vaccine) between study groups, to ensure consistency within the network of trials we analysed. The most relevant or inclusive data (ie, fitting the search criteria and most standard procedures) for a given study, with no differentiation between vaccine manufacturer, were chosen. We follow the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines to report our network meta-analysis.

We curated the retrieved studies on the basis of two predefined outcomes. The first outcome was seroconversion against poliovirus serotypes 1, 2, and 3, measured 4 weeks after the most recent vaccine dose. Seroconversion was defined as a change in antibody titre from non-detectable (<1:8) to detectable (≥1:8) antibody titre, or four-fold or higher increase in antibody titre over the expected decline of maternally derived antibodies (assuming a maternal antibody half-life of 28 days). The second outcome was development of intestinal immunity against serotype 2. Immunity was measured as the absence of shedding of type 2 poliovirus 7 days after a challenge dose of OPV containing the Sabin type 2 strain.

The studies were reviewed and the data extracted independently by two of the investigators (GM and KMO’R). The number of individuals in each study group was recorded by serotype and time of sample collection. We also extracted data for study location, age at administration, route of administration, vaccine antigen content, and challenge vaccine and timing (where applicable). We assessed the risk of bias in accordance with the Cochrane Collaboration’s tool for assessing risk of bias in randomised trials, for individual elements from five domains (selection, performance, attrition, reporting, and other bias) and the overall quality of evidence using the Grading of Recommendations Assessment, Development, and Evaluation framework.

Data analysis

We did a random-effect meta-analysis of single proportions, using an inverse variance pooling method and logit transformation, in the meta package in R (version 3.4.3). A random-effect network meta-analysis was developed for each outcome, with a binomial likelihood and log-link function and computed in a Bayesian framework using the GeMTC package in R (version 3.4.3). Markov chain Monte Carlo (MCMC) simulations estimated posterior distributions of relative treatment effects and SDs, with vague uniform priors. Four independent Markov chains were run with 10000 burn-in iterations and

Figure 1: Systematic review profile

Figure 2: Networks of eligible comparisons for vaccination schedules inducing humoral immunity against poliovirus serotypes 1, 2, and 3
60 000 inference iterations per chain. Convergence of Markov chains was evaluated using the Gelman–Rubin–Brooke diagnostic and time-series plots. Autocorrelation plots were assessed to detect autocorrelation in the chains. Additional analysis included network meta-regression to investigate the effect of study-level covariates, including the estimated mortality rate for children younger than 5 years due to diarrhoeal disease in the country of study location.

We report the pooled random-effect estimates for single-proportion meta-analyses by study arm with 95% CIs. For the network meta-analysis, the between-intervention effect tables between treatments (appendix p 21). Pairwise analysis is additionally reported for programmaticaly important comparisons (appendix p 27).

Model-fit was measured by deviance information criterion, residual deviance, and leverage. We show the SD of the random-effects model (known as $\tau$), as a measure of heterogeneity in the network, where $\tau^2$ is the between-study variance of the true effect size. A node-splitting model was generated to assess inconsistency within the network.

### Role of the funding source
The funders had no role in study design, data collection, data analysis, data interpretation, or writing of the report. The corresponding author had full access to all the data in the study and had final responsibility for the decision to submit for publication.

### Results
A literature search in MEDLINE and CENTRAL identified 437 unique studies and 53 were retrieved for full-text assessment (figure 1). 17 studies describing 47 study groups met the inclusion criteria and were included in the network meta-analysis for humoral immunity, and eight studies describing 25 study groups did so for intestinal immunity (appendix p 6–14). We found a low-to-moderate risk of bias for individual studies and a moderate-to-high quality of evidence for each outcome (appendix p 14). Seven vaccine formulations were included in the total analysis: tOPV, bOPV, Salk IPV (referred to as IPV and administered intramuscularly), mIPV2 (administered intramuscularly), fIPV (one-fifth Salk IPV dose, administered intradermally), sIPV (administered intramuscularly), and aluminium-adjuvanted IPV (IPV-Al; one-tenth Salk IPV dose, administered intramuscularly).

There were 19 unique vaccination schedules identified, and each formed a node in the network (figure 2) with corresponding nodes by patient age (table). The network was separated into two subgroups, studies without a birth dose (14 nodes, 15 studies) and those with a birth dose (five nodes, two studies). In total, 8254 infants were included in the analysis for seroconversion against serotype 1, 8241 against serotype 2, and 8279 against serotype 3. The pooled proportion for all three serotypes corresponded schedules by patient age (table). The network was separated into two subgroups, studies without a birth dose (14 nodes, 15 studies) and those with a birth dose (five nodes, two studies). In total, 8254 infants were included in the analysis for seroconversion against serotype 1, 8241 against serotype 2, and 8279 against serotype 3. The pooled proportion for all three serotypes ranged from 13% to 100% between different vaccine schedules.

In the network meta-analysis, the criteria for adequate model-fit were satisfied with no significant inconsistency between any direct and indirect estimates ($p<0.05$ for all comparisons) and low between-trial heterogeneity ($t=0.05$, 95% CrI 0.009–0.15; appendix p 18).

All vaccine schedules achieved higher seroconversion RRs against serotype 2 than the bOPV alone (figure 3). The addition of a single dose of IPV to a three-dose bOPV schedule significantly improved seroconversion (at age 18 weeks; RR 0.85, 95% CrI 0.75–1.00), and a second incremental increase in RR was seen after a second dose of IPV (at 36 weeks; 1.00–1.00–1.00). A pairwise analysis of studies that directly compared seroconversion from three bOPV doses plus two IPV doses with three bOPV doses plus one IPV doses gave a pooled RR of 1.25 (95% CrI 1.07–1.47; appendix p 27).

Combined bOPV-IPV schedules where IPV administration preceded bOPV had a lower RR than in cases in which IPV succeeded bOPV (figure 3).

| Age schedule (weeks) |   |
|----------------------|--|
| IPV + 2 bOPV        | 8, 12, 16 or 8, 16, 24 |
| 2 fIPV*             | 6, 14 |
| 2 IPV               | 6, 14 |
| 2 fIPV + bOPV       | 8, 12, 16 or 8, 16, 24 |
| 3 bOPV              | 6, 10, 14 |
| 3 bOPV + IPV        | 6, 10, 14, 14 |
| 3 bOPV + mIPV2      | 6, 10, 14, 14, 36 |
| 3 bOPV + 2 IPV      | 6, 10, 14, 14, 36 |
| 3 fIPV*             | 6, 10, 14 or 8, 16, 24 |
| 3 IPV               | 6, 10, 14, 8, 12, 16, or 8, 16, 24 |
| 3 IPV-Al            | 6, 10, 14 |
| 3 sIPV             | 8, 12, 16 |
| 3 sIPV              | 8, 10, 14 or 8, 12, 16 |
| 4 bOPV + IPV        | Birth, 6, 10, 14 |
| 4 bOPV + sIPV       | Birth, 6, 10, 14, 14 |
| 4 bOPV + 2 IPV      | Birth, 6, 10, 14, 18 |
| 4 IPV + mIPV2       | Birth, 6, 10, 14 |
| 4 tOPV              | Birth, 4, 8, 12 or birth, 6, 10, 14 |
| fIPV* + bOPV + fIPV*| 6, 10, 14 |

bOPV=bivalent oral poliovirus vaccine. fIPV= fractional inactivated poliovirus vaccine. IPV=conventional inactivated poliovirus vaccine. IPV-Al= aluminium hydroxide adjuvanted inactivated poliovirus vaccine. mIPV2= monovalent serotype 2 high-dose inactivated poliovirus vaccine. sIPV= Sabin inactivated poliovirus vaccine. tOPV= trivalent oral poliovirus vaccine. ‘One-fifth full IPV dose administered intradermally: 15D antigen content for serotype 2 administered intramuscularly. One-tenth full IPV dose administered intramuscularly. 3S: 6, 32, and 45 D antigen content for serotype 1, 2 and 3, respectively, administered intramuscularly.

Table: Vaccine schedules by approximate age of administration included for each node of the network meta-analysis for humoral immunity.
We compared IPV-only schedules for both two and three doses (see appendix p 21 for a table of relative effects). There was no significant difference between the seroconversion achieved by two doses of full-dose IPV and intradiscal fIPV (RR 0.96, 0.93–1.00). Additionally, there was little difference between two doses of fIPV and three doses of full-dose IPV, and 1.01 (0.85–1.20) for the network meta-analysis of serotype 1. There was high between-trial heterogeneity for serotype 3. There was no significant difference between three doses of bOPV alone and three doses of fIPV + bOPV + fIPV, 0.80 (0.68–0.93) for sIPV or IPV-Al, or between two doses of IPV and fIPV (RR 0.88, 0.78–0.98).

For schedules including a birth dose from the full network analysis, the addition of one or two doses of IPV following a four-dose bOPV-only schedule increased seroconversion (figure 3).

For the network meta-analysis of serotype 1, there was significant inconsistency between direct and indirect effects in one network comparison, and in six comparisons for serotype 3. There was high between-trial heterogeneity for both serotypes (serotype 1 τ=0.05, 95% CrI 0.009–0.15) and (B) those with a birth dose (0.23, 0.026–1.31). Between-trial heterogeneity for (A) schedules without a birth dose (τ=0.05, 0.09–0.32; appendix p 20).

There were 15 unique vaccination schedules identified for two versus three doses of intradiscal fIPV. Additional, there was no significant difference between three doses of bOPV alone and three doses of fIPV + bOPV + fIPV, 0.80 (0.68–0.93) for sIPV or IPV-Al, or between two doses of IPV and fIPV (RR 0.88, 0.78–0.98).

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Network meta-analysis of routine immunisation schedules for intestinal immunity against serotype 2

Figure 5: Network meta-analysis of routine immunisation schedules for intestinal immunity against serotype 2

(A) Network plot. Each node represents a vaccine schedule and the lines are direct comparisons. The width of the lines is proportional to the number of trials compared and the node size is proportional to the number of trials that include direct comparisons. The width of the lines is proportional to the number of trials compared and the node size is proportional to the number of trials that include direct comparisons. The width of the lines is proportional to the number of trials compared and the node size is proportional to the number of trials that include direct comparisons. The width of the lines is proportional to the number of trials compared and the node size is proportional to the number of trials that include direct comparisons.

Discussion

To our knowledge, we report the first application of a network meta-analysis to assess the immunogenicity of vaccine schedules against poliomyelitis and provide a single, comprehensive analysis of polio routine immunisation schedules for humoral and intestinal outcomes. We found that for humoral immunity the addition of one dose of IPV to bOPV schedules increases the immunity against serotype 2, that there is no difference in relative immunogenicity of IPV variants (Salk IPV, sIPV, intradermal fIPV, IPV-Al), and that the timing of the IPV dose in bOPV-IPV schedules is associated with immunogenicity. Potentially, the most important finding of our study is for mucosal immunity, as we found no evidence of increased intestinal immunity against serotype 2 associated with an addition of IPV to a bOPV-only schedule.

Clinical trials have provided valuable information to inform policy but are limited by specific head-to-head schedule comparisons, small cohort sizes, and generation of country-specific data. Previous literature reviews and meta-analyses have addressed specific aspects of vaccination against poliomyelitis virus: mucosal immunity induced by OPV versus IPV, humoral immunity of one versus two doses of IPV, and immunogenicity of two doses of fIPV. Additionally, reviews and meta-analysis have compared the immunogenicity of bOPV-IPV mixed schedules and IPV alone. In agreement with our results, a meta-analysis by Tang and colleagues found no significant difference between IPV-only and IPV-OPV schedules in seroconversion against serotypes 1 and 3. However, they included only six trials with two different schedule groups (IPV-only and IPV-OPV mixed schedule) and provided no data on mucosal immunity. Therefore, our analysis goes beyond what previous studies have done.

After the switch from tOPV to bOPV, SAGE recommended the introduction of at least one dose of IPV at age 14 weeks or older to provide an immunity base to type 2 poliovirus. Our results provided evidence that individuals vaccinated with bOPV-only schedules have negligible immunity against poliovirus 2 (probably from passive type-2 exposure or antibody cross-neutralisation from types 1 and 3). This immunity deficit suggests that the estimated 43 million children across 33 countries who did not receive IPV because of supply shortages have no protection. Notably, the addition of a single dose of IPV (at 14 weeks) improved humoral immunity against serotype 2, whereas a second dose (at 18 or 36 weeks) had a smaller impact and a single mIPV2 dose provided equivalent immunogenicity to two doses of trivalent IPV. Our results also highlight that the order and timing of the IPV dose in mixed schedules is important, with a reduced immunogenicity against the type 2 strain in cases where IPV preceded bOPV. This effect is likely due to the earlier age at which IPV is administered and the effect of maternal antibodies.
Of note, the addition of IPV had a negligible effect on the development of intestinal immunity against serotype 2. Although the ability of IPV to induce humoral immunity is established, IPV has a more complex role in mucosal immunity. IPV-only schedules provide inadequate intestinal immunity and do not prevent viral shedding following a challenge dose, but they might reduce the quantity and duration of shedding.6,27 However, IPV has been shown to boost mucosal protection in OPV-primed individuals.6,27 Our results provide evidence that the prime-boost model established for IPV works in a serotype-specific manner.

Our findings have several limitations. Consistency of the network is a fundamental assumption of network meta-analyses,8,9 which was not met for serotypes 1 and 3 for which heterogeneity and inconsistency persisted through the subgroup and regression analyses (appendix p 18–20). The type, schedule, and immunogenicity of poliovirus vaccines vary by location.1 The studies in this analysis1 were done in eastern Mediterranean and Latin American countries that have primary vaccine schedules in which the first (non-birth) dose is administered between 4 and 8 weeks. Therefore, our results are primarily useful for policy makers in these settings. The geographical and age-schedule variation in absolute immunogenicity is incorporated into our study as the network meta-analysis method models the relative effects between vaccines, thus eliminating differences in baseline immunogenicity of comparator schedules.8,9

The most widely adopted measure of poliovirus mucosal immunity is through administration of a challenge dose of OPV and a subsequent collection of stool samples. Extrapolating intestinal immunity and transmission impact on the basis of absence of shedding 7 days following a challenge dose does not capture the duration of shedding, quantity of virus shed, or nasopharyngeal immunity.8,9,25 However, these measures represent the best proxy for intestinal immunity to poliovirus available from clinical trial data. Finally, our analysis only provides estimates on protection within the timescale of the trials.

There are research gaps highlighted in our modelled networks, particularly the need for evaluation of mucosal immunity in more studies. A schedule of three bOPV doses followed by fIPV has not been included in randomised controlled trials, yet this schedule has been adopted in India and Sri Lanka.9 Research is needed to compare IPV-only vaccine schedules in a post-eradication setting and address the need to develop a live vaccine with improved genetic stability. A novel OPV, which is a live attenuated vaccine with a lower risk of reversion than the standard OPV, and an IPV plus dmLT adjuvant are being trialled for induction of mucosal immunity.1,25-27

The findings of our comprehensive analysis show that a network meta-analysis is effective to evaluate multi-arm vaccination studies. Our results support, with policy recommendations from the SAGE, the addition of IPV into routine immunisation schedules and the adoption of affordable IPV approaches. We show that a single dose of IPV improves humoral immunity against serotype 2 and suggest that in times of IPV supply constraints, equitable distribution of a single dose of IPV should be prioritised over cohorts receiving a second dose, taking into account country-specific risk. This IPV addition will be unlikely to prevent faecal–oral transmission of the virus, but would provide individual protection against paralytic disease.

Contributors
GRM, WJE, and RWS contributed to study concept and design. GRM and KMO’R independently reviewed studies and extracted data, did the statistical analysis, and drafted the report. GM, WJE, OM, RWS, NCG, ASB, and KMO’R contributed to critical revision of the manuscript for intellectual content.

Declaration of interests
We declare no competing interests.

Acknowledgments
The funding source for this research was through a UK Medical Research Council PhD studentship held by GRM. We thank the Polio Research Committee for their discussions that initially highlighted the importance of this research.

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