Turkish Validation of the Work-Family Conflict Scale

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ABSTRACT

The aim of this study is to translate and validate work-family conflict scale abbreviated and developed by Matthews, Kath and Barnes-Farrell (2010) in a Turkish sample. The present study was conducted in a sample of 317 bank employees. The confirmatory factor analysis revealed two-factor structure work-to-family and family-to-work conflict dimensions of the original scale. The convergent, discriminant, and nomological validity analyses showed satisfactory results. Reliability analyses indicated that the scale had high internal consistency and sufficient item correlations. The findings revealed that the Turkish version of work-family conflict scale is a valid and reliable instrument to examine work-family conflict in Turkish context.

Keywords: work-family conflict, scale validation, validity, reliability, Turkish sample

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Work and family roles occupy both important places in employees’ lives. These two individual roles in life are not isolated from each other. On the contrary, they are associated in various ways (Ilies, Liu, Liu, & Zheng, 2017). This relationship has widely examined by “role scarcity” and “negative spillover” perspectives in many studies (Tetrick & Buffardi, 2006). As a consequence of these approaches, the most widely studied aspect of work-family interface has been work-family conflict (Greenhaus & Powell, 2003).
Work-family conflict can be defined as “a form of inter role conflict in which the role pressures from the work and family domains are mutually incompatible in some respect” (Greenhaus & Beutell, 1985, p. 77). This conflict can occur from both directions whereas work life can have an effect on family life and family life can have an effect on work life (Tetrick & Buffardi, 2006). Both of these domain interferences have significant effects on work related, family related, health related outcomes, including job satisfaction, career satisfaction, organizational commitment, absenteeism, turnover intentions, job performance, life satisfaction, family satisfaction, burnout, psychological strain, work-related stress, and family-related stress (Anderson, Coffey, & Byerly, 2002; Aycan & Eskin, 2005; Bruck, Allen, & Spector, 2002; Burke & Greenglass, 2001a; Grandey, Cordeiro, & Crouter, 2005; Lapierre, Spector, Allen, Poelmas, Coopere, O’Driscoll, Sanchez, Brough, & Kinnunen, 2008; Michel, Mitchelson, Kotrba, LeBreton, & Baltes, 2009; Nohe & Sonntag, 2014).

There are several measures in order to assess work-family conflict in the literature. Bellavia and Frone (2005), MacDermid (2005), and Tetrick and Buffardi (2006) summarized work-family conflict measures in a detail. Tetrick and Buffardi (2006) stated that early measures of work-family conflict are bidirectional, which assess work life interfering with family life (e.g. Bohen & Viveros-Long, 1981). Carlson, Kacmar, and Williams (2000) noted that following measures have taken into consideration the directionality of work-family conflict meaning work life interfering with family life and as well as family life interfering with work life (e.g. Frone, Russell, & Cooper, 1992; Gutek, Searle, & Klepa, 1991; Kopelman, Greenhaus, & Connoly, 1983, Netemeyer, Boles, & McMurrion, 1996). Recent studies have recognized various sources of conflict and have created distinct dimensions for time-based conflict, strain-based conflict, and behavior-based conflict (e.g. Carlson et al., 2000; Stephens & Sommer, 1996). Even though Stephens and Sommer (1996) utilized time-based, strain-based, and behavior-based conflict dimensions, they only measured work-family conflict, which refers to work life interfering with family life. On the other hand, Carlson et al. (2000) developed a multidimensional measure containing the three types and two directions of work-family conflict which leads to time-based work-to-family conflict, strain-based work-to-family conflict, behavior-based work-to-family conflict, time-based family-to-work conflict, strain-based family-to-work conflict, and behavior-based family-to-work conflict subscales.

However, as Matthews, Kath, and Barnes-Farrell (2010) stated, elongated measures can be an issue in some studies such as longitudinal research designs whereas a concise instrument containing theoretical relevant items proves to be more advantageous. The aim of the present study is to translate and validate this measure by Turkish sample and utilize it in the Turkish context. Thereby it is expected to contribute to the literature by enabling this instrument on the work-family and family-work conflict, which can have important consequences for both the employees work life and family life as well as for the organization.

Background
The interaction between work life and family life has received much interest over past several decades especially due to the changes in socioeconomic and family structures such as rising number of women in the workforce, increasing working hours, the grate share of dual-earner families, the growing rate of working single parents, and responsibility of child care and aging
parents (Annor, 2013; Kinnunen, Rantanen, Mauno, & Peeters, 2014; Mokomane, 2013; Parasuraman & Greenhaus, 1997). Moreover, globalization of economy and technological developments changed work life and working conditions by giving means of working out of the office remotely and after office hours (Kinnunen et al., 2014) which eventually leads in the interaction between work roles and family roles. This interaction can be either positive or negative (Allen, 2013; Carlson & Grzywacz, 2008; Kinnunen et al., 2014). Negative work-family interaction in other words work-family conflict has been the most widely studied subject in the work-family literature (Poelmans, O’Driscoll, & Beham, 2005).

The majority of the literature based theoretical foundation of the work-family conflict on Kahn, Wolfe, Quinn, Snoek, and Rosenthal’s (1964) role stress theory and the scarcity approach (Allen, 2013; Grzywacz & Butler, 2008; Kinnunen et al., 2014; MacDermid, 2005). According to role theory, taking part in one role becomes incompatible with taking part in the other role resulting in role conflict (Greenhaus & Beutell, 1985). Scarcity approach suggests that people have limited resources such as time, energy, and emotions and these individual reserves can easily drain while participating these various roles (Goode, 1960; Marks, 1977). As both roles can interfere with one another, work roles interact with family roles (work-to-family conflict), and family roles interact with work roles (family-to-work conflict) conflict can occur from both domains indicating that work-family conflict is a bidirectional concept (Allen, 2013; Frone, 2003; MacDermid, 2005).

Greenhaus and Beutell (1985) suggested that the type of conflict between work and family roles could have three major bases, time-based conflict, strain-based conflict, and behavior-based conflict involving both domains. Time-based conflict emerges when time required for one role obstructs participating in the other role. Strain-based conflict occurs when the strain caused by one role prevents fulfilling the demands of the other role. Behavior-based conflict arises when behaviors appropriate for one role are not suitable for the expected behaviors of the other role (Allen, 2013; Geurts & Demerouti, 2003; Greenhaus & Beutell, 1985; Kinnunen et al., 2014; MacDermid, 2005; Matthews et al., 2010; Parasuraman & Greenhaus, 1997).

Work-to-family conflict and family-to-work conflict have significant effects on the organizations as well as on the employees and their families. As Kinnunen et al (2014) pointed out early studies suggested that work-to-family conflict has effects on the family area whereas family-to-work conflict has consequences for the work area implying that each domain has outcomes for the cross domain. However, in their meta-analysis Allen, Herst, Bruck, and Sutton (2000) revealed that work-to-family conflict has a correlation with not only non-work-related outcomes but also work-related and health-related outcomes in contrast to the cross-domain rule. This implies that both domains could have effects on work-related, non-work-related, and health-related consequences.

Work-family conflict has negative effects on individuals’ health and well-being (Bellavia & Frone, 2005; Geurts & Demerouti, 2003). O’Driscoll, Brouh, and Kalliath (2004) revealed in their study that both work-family conflict domains are associated with physical health symptoms and psychological strain. Van Steenbergen and Ellemers (2009) found that work-to-family conflict has negative effects on physical health while family-to-work conflict has no significant effect. Several other studies reported similar results regarding physical health.
outcomes and psychological strain (Gareis, Barnett, Ertel, & Berkman, 2009; Grzywacz, 2000; Kalliath, Kalliath, & Chan, 2014).

Frone’s (2000) research results indicated that work-to-family and family-to-work conflict linked to having a mood, anxiety, and substance abuse disorders. Frone, Russell and Cooper (1997) revealed that conflict between work and family roles in both domains lead to depression and heavy alcohol use. In their meta-analysis, Amstad, Meier, Fasel, Elfering, and Semmer (2011) reported that work interference with family and family interference with work are both related to anxiety, depression, stress, and substance abuse. Similarly, Wang and Peng (2017) reported positive relationships between work-to-family conflict, family-to-work conflict and depression.

In their study, Burke and Greenglass (2001b) found that family-to-work conflict has positive associations with burnout. Similarly, Bacharach, Bamberger, and Conley (1991) reported positive relationships between work-to-family conflict and burnout in two different samples. Rupert, Stevanovic, and Hunley (2009) found that both work-family conflict domains are related to burnout subscales, personal accomplishment, emotional exhaustion, and depersonalization.

Regarding the effects of the work-family conflict on satisfaction such as job satisfaction, life satisfaction, and family satisfaction, Netemeyer et al. (1996) showed in their study that work-to-family conflict and family-to-work conflict are in association with job satisfaction, life satisfaction, and family satisfaction negatively in three different samples. Moreover, O’Driscoll et al. (2004) revealed that both domains are negatively related with job satisfaction and family satisfaction. In their meta-analysis, Kossek and Ozeki (1998) found that work-to-family conflict and family-to-work conflict have negative correlations with job satisfaction and life satisfaction. Shockley and Singla (2011) showed in the meta-analysis that work interfering with family and family interfering with work are correlated negatively with job satisfaction and family satisfaction. In their more recent meta-analysis, Amstad et al. (2011) reported that work-family-conflict in both domains are inversely related to life satisfaction, marital satisfaction, and family satisfaction.

Furthermore, Aryee, Srinivas, and Tan (2005) demonstrated that work-family conflict in both domains are related to organizational commitment in a negative way. In addition to this, Netemeyer et al. (1996) found similar results that work-family-conflict is associated with organizational commitment in opposite directions. Moreover, Amstad et al. (2011) revealed in their meta-analysis that higher levels of work interference with family and family interference with work is correlated at lower levels of organizational commitment.

With respect to the effect of work-to-family conflict and family-to-work conflict on job performance and family performance, Frone, Yardley, and Markel (1997) found that work-family-conflict in both domains are negatively related to job performance and family performance. In addition to this, in their meta-analysis, Hobler, Hu, and Wilson (2010) reported that higher levels of work-family-conflict is associated with lower levels of self-rated and manager-rated job performance.

Moreover, in regard to the negative influence of work-family-conflict on withdrawal behavior, absenteeism, and turnover intentions, in their meta-analysis, Mesmer-Magnus and Viswesvaran (2005) reported that work-family conflict and family-to-work both are
correlated to withdrawal from work. Anderson et al. (2002) and Boyar, Maertz, and Pearson (2005) revealed that higher levels of work-family conflict in both directions are associated with high levels of absenteeism. Furthermore, several studies showed that work-family conflict in both domains are related with high levels of turnover intentions (Anderson et al., 2002; Boyar, Maertz, Pearson, & Keough, 2003; Greenhaus, Parasuraman, & Collins, 2001; Netemeyer, Brashear-Alejandro, & Boles, 2004).

**Method**

**Translation Process**

In order to translate the work-family conflict scale, the method developed by Bristlin, Lonner, and Thorndike (1973) was utilized. Initially, the scale was translated into Turkish from the original language by two language experts who had experience in English speaking countries. Secondly, two professors from the field of organizational behavior assessed the scale. In the next step, the experts back translated the Turkish version into original language and this version was examined in detail. The two professors checked the final version of the work-family conflict scale at the last step of the translation process.

**Sample**

The sample in this study comprised of bank employees in Adana, Turkey. In total, 350 questionnaires were distributed, 330 of them were returned. Due to several reasons, (e.g. missing information) 13 questionnaires were eliminated, and thus the sample consisted of 317 employees in total. The sample included 48.9% (155) female. The participants mean age was 37.96 years (SD = 6.73) and the mean job tenure was 7.01 years (SD = 5.35). Most of the participants had a bachelor’s degree (81.1%) regarding their education level.

**Instruments**

**Work-Family Conflict**

Work-Family conflict was measured with using two-dimensional scale abbreviated and developed by Matthews, Kath and Barnes-Farrell (2010) using the long version of the scale of work–family conflict by Carlson, Kacmar, and Williams (2000). This scale consists of six items which assesses two factors, work-to-family and family-to-work conflict. 5-point Likert scale was used (1 = strongly disagree, 5 = strongly agree) to assess the scale. Matthews et al. (2010) reported alpha coefficients for work-to-family conflict dimension ranging from .75 to .80 whereas for family-to-work conflict dimension ranging from .71 to .72. In the present study, Cronbach’s alpha and McDonald’s omega reliability coefficients for work-to-family conflict factor were .80 and .81 whereas these coefficients for family-to-work conflict factor were .77 and .78.

**Life Satisfaction**

Life satisfaction was assessed using a 5-item Life Satisfaction scale created by Diener, Emmons and Griffin (1985). The original scale rated on a 7-point Likert scale, but this study used a 5-point scale (1 = strongly disagree, 5 = strongly agree). Diener et al. (1985) and
reported Cronbach’s alpha value of .87. In this study, both Cronbach’s alpha and McDonald’s omega coefficients were found to be .89.

Subjective Happiness
Subjective happiness scale (Lyubomirsky & Lepper, 1999) was used regarding well-being. This scale consists of 4-items, which uses 7-point Likert scale. In this study, 5-point frequency scale was employed. Lyubomirsky and Lepper (1999) obtained a Cronbach’s alpha value of .87. In the present study, Cronbach’s alpha and McDonald’s omega were found to be .87 and .88.

Depression
Depression was examined using the depression dimension of work-related depression, anxiety, and irritation scale, which was developed by Caplan, Cobb, French, Van Harrison, and Pinneau (1980). This measure has 6-items and uses a 5-point Likert scale (1 = never, 5 = always) in the present study. The reliability coefficient of this scale ranged from .81 to .88 in several studies (Begley & Czajka, 1993; Jalajas, 1994; Kaufmann & Beehr, 1986). In this study, alpha was found to be .92 and omega was .93.

Statistical Analyses
The two-factor model developed by Matthews et al. (2010) was tested. Data analyses were performed with R Studio version 0.99.903 based on R Version 3.3.1 (R Core Team, 2016). In addition to R core packages psych (Revelle, 2017), Qgraph (Epskamp, Cramer, Waldorp, Schmittmann, & Borsboom, 2012), and SemPlot (Epskamp with Stuber, 2017) were also utilized for data analyses purposes. Confirmatory factor analysis was conducted using Lavaan package (Rosseel, 2012) with maximum likelihood estimation method. Model fit was assessed by several goodness-of-fit statistics, the chi-square ($\chi^2$) statistic, the comparative fit index (CFI; Bentler, 1990), the Tucker-Lewis fit index (TLI; Tucker & Lewis, 1973), the root mean square error of approximation (RMSEA; Steiger & Lind, 1980), and the standardized root mean square residual (SRMR; Bentler, 1995). For assessing the construct validity, convergent validity, discriminant validity, and nomological validity were examined (Hair, Black, Babin, & Anderson, 2014, pp. 631). According to Hair et al. (2014, pp. 605) to establish convergent validity, the standardized factor loadings should be .50 or higher and optimally .70. In addition to this, the average variance extracted (AVE) values should be greater than .50 whereas composite reliability coefficients (CR) should be .70 or higher for the convergent validity and for internal consistency (Hair et al., 2014, pp. 605). To assess discriminant validity, maximum shared variance (MSV) and average shared variance (ASV) should be lower than average variance extracted (AVE) and moreover the squared root of AVE’s should be higher than the correlations between the variables (Hair et al., 2014, pp. 631). To examine nomological validity, the assumed relationships between constructs are evaluated with correlation analysis. For assessing the internal consistency of the scales, Cronbach’s alpha (Cronbach, 1951) and McDonald’s omega (McDonald, 1999) coefficients (Composite reliability coefficient) were used. Moreover, item and item-total statistics were analyzed to examine the reliability of the scales. According to Nunnally and Bernstein (1994, pp. 305), the corrected item-total correlations should be .30 or higher.
Results

Reliability Results

For assessing the reliability, both item analysis and internal consistency were examined. As shown in Table 1, item-total correlations were obtained between .59 and .72. This result meets the requirement of item-total correlations as being higher than .30 (Nunnally & Bernstein, 1994, pp. 305). In addition to this, item analysis revealed that deleting any items from the scale would not increase Cronbach’s alpha values.

Table 1

| Items                  | Corrected Item-Total Correlation | Alpha If Item Deleted | M    | SD    |
|------------------------|----------------------------------|-----------------------|------|-------|
| Work-to-Family Conflict Items |
| WF1                    | .72                              | .65                   | 1.57 | .74   |
| WF2                    | .59                              | .79                   | 1.59 | .75   |
| WF3                    | .63                              | .74                   | 1.50 | .74   |
| Family-to-Work Conflict Items |
| FW1                    | .62                              | .68                   | 1.62 | .75   |
| FW2                    | .62                              | .68                   | 1.63 | .76   |
| FW3                    | .59                              | .72                   | 1.33 | .60   |

Internal consistency of the scales was analyzed using both Cronbach’s alpha and McDonald’s omega coefficients. According to the results (Table 4) alpha coefficient for work-family conflict dimension is .80 and omega coefficient is .81 whereas alpha coefficient for family-work conflict dimension is .77 and omega coefficient is .78. These internal consistency values which are above .70 show acceptable results.

Confirmatory Factor Analysis (CFA)

The confirmatory factor analysis results show very good model fit (Table 2). According to model fit indices, chi-square is 9.54 with 8 degrees of freedom and p-value is .298, which is statistically not significant. Moreover, CFI and TLI have a value of .99 whereas SRMR is .02 and RMSEA has a value of .02.

Table 2

| Fit Index   | Model Results | Reference Values |
|-------------|---------------|------------------|
| $\chi^2$    | 9.54          | $0 \leq \chi^2 \leq 2df$ |
| df          | 8             | -                |
| $\chi^2$ / df | 1.19          | $0 \leq \chi^2$/df $\leq 2$ |
| p-value     | .29           | $P > .00$        |
| CFI         | .99           | .90 $< CFI$      |
| TLI         | .99           | .90 $< TLI$      |
| RMSEA       | .02           | .00 $< RMSEA < .05$ |
| SRMR        | .02           | .00 $< SRMR < .05$ |

According to confirmatory factor analysis, the standardized factor loadings are given in Table 3. These factor loadings range from .67 to .87 showing the model is well defined. The
correlation between these factors, work-family conflict and family-work conflict is .46 as seen in Figure 1.

Table 3
Factor Loadings of the Work-Family Conflict Scale

| Items                  | Work-to-Family Conflict | Family-to-Work Conflict |
|------------------------|-------------------------|-------------------------|
| Work-to-Family Conflict Items |
| WF1                    | .87                     |                         |
| WF2                    | .67                     |                         |
| WF3                    | .74                     |                         |
| Family-to-Work Conflict Items |
| FW1                    | .76                     |                         |
| FW2                    | .74                     |                         |
| FW3                    | .69                     |                         |

Figure 1. Work-family conflict scale confirmatory factor analysis results

Validity Results
After establishing factorial validity with confirmatory factor analysis, for further evaluation of construct validity, convergent validity, discriminant validity, and nomological validity have been examined. For assessing convergent validity of work-family conflict scale, several criteria were utilized. Firstly, factor loadings should be higher than .50 or preferably .70 (Hair et al., 2014, pp. 605). According to the analysis results as shown in Table 3, the minimum factor loading is .67 which ensures convergent validity. Secondly the average variance extracted values should be higher than .50 for establishing the convergent validity (Hair et al., 2014, pp. 605). According to the results (Table 4), average variance extracted value for work-to-family conflict is .59 and for family-to-work conflict is .54 which are both above .50. In addition, the last criterion for convergent validity is providing composite reliability coefficients higher than .70 (Hair et al., 2014, pp. 605) and according to the results (Table 4), composite reliability coefficient for work-to-family conflict is .81 and for family-to-work conflict is .78. These three rules suggest that convergent validity is established for the work-family conflict scale.
Table 4
Means, Standard Deviations, Cronbach’s Alphas, McDonald’s Omegas, Composite Reliabilities, AVEs, and Correlations of the Scales Used

| Variables | α    | CR-ω | AVE | MSV | ASV | Mean | SD  | WFC | FWC | LS | SH | D      |
|-----------|------|------|-----|-----|-----|------|-----|-----|-----|----|----|-------|
| WFC       | .80  | .59  | .22 | .19 | .155| .63  | .77 |     |     |    |    |       |
| FWC       | .77  | .54  |     | .30 | .153| .58  | .47**| .73 |     |    |    |       |
| LS        | .89  | .61  | .33 | .30 | .445| .60  |     | .43**| .58**| .78|    |       |
| SH        | .87  | .64  | .38 | .31 | 4.04 | .72  |     | .42**| .58**| .80|    |       |
| D         | .92  | .68  | .31 | 1.95| .75  | .43**| .55**| .59**| .62**| .82|    |       |

WFC=Work to Family Conflict, FWC=Family to Work Conflict, LS=Life Satisfaction, SH=Subjective Happiness, D=Depression, α=Cronbach’s Alpha Coefficient, CR-ω=Composite Reliability, McDonald’s Omega Coefficient, AVE=Average Variance Extracted, MSV=Maximum Shared Variance, ASV=Average Shared Variance, SD=Standard Deviation, *Diagonal values are the square root of AVEs, **p<0.01.

In order to examine the discriminant validity of the scale, the principles concerning AVE values were taken into consideration. According to these principles, AVE values should be higher than MSV and ASV values, whereas the square root of AVE values should be higher than the correlations between the variables (Hair et al., 2014, pp. 631). As displayed in Table 4, the AVE value for work-to-family conflict is .59 and for family-to-work conflict is .54 while MSV values and ASV values are lower than these results. In addition to this, in Table 4, the diagonal values show the square root of AVEs that are higher than the correlations between the variables. These findings reveal that discriminant validity is established for the work-family conflict scale.

To assess the nomological validity, the relationships between the variables were examined. As shown in Table 4, the correlations of work-to-family conflict and family-to-work conflict with life satisfaction and subjective happiness are positive whereas with depression are negative as expected according to the theoretical framework. These findings suggest that work-family conflict scale has satisfactory reliability and construct validity comprising convergent, discriminant, and nomological validities.

Discussion

The aim of this study was to examine the validity and reliability of the work-family conflict scale developed by Matthews et al. (2010) in a Turkish sample. The results of this research show satisfactory support for both validity and reliability of the work-family conflict scale in Turkish sample. Confirmatory factor analysis results produce two-factor structure work-to-family conflict and family-to-work conflict as in the original scale. Factor loadings were above .50 mostly above .70 and were all statistically significant. These findings show consistency with the earlier studies conducted in various cultures and contexts (Annor & Amponsah-Tawiah, 2017; Hill, Morganson, Matthews, & Atkinson, 2016; Rudolph, Michel, Harari & Stout, 2014).

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In addition to this, the study established construct validity of the two-factor scale by showing convergent validity, discriminant validity, and nomological validity. This result is consistent with previous studies (Annor & Amponsah-Tawiah, 2017; Matthews et al., 2010).

With regard to reliability of the scale, the results showed sufficient internal consistency and reliability. Cronbach’s alpha and McDonald’s omega coefficients for both subscales were obtained above .70. Cronbach’s alpha values for work-to-family conflict was .80 whereas for family-to-work conflict was .77. McDonald’s omega coefficient for work-to-family conflict was .81 and for family-to-work conflict was .78. In addition to these values, item-total correlations, which are another indicator of reliability, showed satisfactory results between .59 and .72. These reliability values are consistent with the result conducted previously (Annor & Amponsah-Tawiah, 2017; Hill et al. 2016; Matthews et al., 2010; Rudolph et al., 2014).

As Matthews et al. (2010) suggested that the abbreviated version of the work-family conflict scale is especially useful for studies requiring fewer items in the survey and longitudinal studies. This short version of work-family conflict measure is convenient in order to examine the overall aspect of work-family conflict rather than assessing the different basis such as time-based, strain-based, or behavior-based work-family and family-work conflict dimensions and theoretical distinctions. Therefore, the long version developed by Carlson et al. (2000) is more relevant as pointed out by Matthews et al. (2010).

This study has several limitations. The research was conducted only in one sample due to the financial and time constraints. Regarding the participants, only bank employees were recruited which can have an effect on the behaviors encountered. Moreover, the nomological validity was assessed with only three variables. In the future, the studies may conduct in multiple samples, with participants from several occupations, and examine various outcome and predictor variables in relation with work-family conflict.

In future studies, this valid and concise scale can be used to examine the relationships between work-family conflict and several outcomes important for family life, work life, and personal well-being of employees. In addition to this, regarding the concise nature of the scale, future research may conveniently use in longitudinal designs in order to assess work-family conflict, its outcomes, and predictors.

The present study indicates that the work-family conflict scale has two-factor structure, and these dimensions show sufficient reliabilities. In addition to these properties, this study shows that the construct validity involving convergent validity, discriminant validity, and nomological validity were established for the Turkish sample. These results indicate that work-family conflict scale shows satisfactory psychometric characteristics and this scale is a valid and reliable construct to use in studies for Turkish samples.

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