Non-Gaussianity of low frequency heart rate variability and sympathetic activation: lack of increases in multiple system atrophy and Parkinson disease

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INTRODUCTION
The correlates of indices of long-term ambulatory heart rate variability (HRV) of the autonomic nervous system have not been completely understood. In this study, we evaluated conventional HRV indices, obtained from the daytime (12:00–18:00) Holter recording, and a recently proposed non-Gaussianity index (λ; Kiyono et al., 2008) in 12 patients with multiple system atrophy (MSA) and 10 patients with Parkinson disease (PD), known to have varying degrees of cardiac vagal and sympathetic dysfunction. Compared with the age-matched healthy control group, the MSA patients showed significantly decreased HRV, most probably reflecting impaired vagal heart rate control, but the PD patients did not show such reduced variability. In both MSA and PD patients, the low-to-high frequency (LF/HF) ratio and the short-term fractal exponent α1, suggested to reflect the sympathovagal balance, were significantly decreased, as observed in congestive heart failure (CHF) patients with sympathetic overdrive. In contrast, the analysis of the non-Gaussianity index λ showed that a marked increase in intermittent and non-Gaussian HRV observed in the CHF patients was not observed in the MSA and PD patients with sympathetic dysfunction. These findings provide additional evidence for the relation between the non-Gaussian intermittency of HRV and increased sympathetic activity.

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In the present study, this conjecture is more directly tested by studying long-term ambulatory HRV in patients with multiple system atrophy (MSA). MSA is a sporadic and rapidly progressive neurodegenerative disorder that presents with autonomic failure in combination with Parkinsonism or cerebellar ataxia (Wenning et al., 2008; Stefanova et al., 2009).

The autonomic symptoms are believed to be due to neuropathological abnormalities in both preganglionic sympathetic (Sone et al., 2005) and vagal (Benarroch et al., 2006) neurons. In previous HRV studies, decreased high frequency, vagally mediated HRV was observed in MSA patients than in age-matched healthy controls (Gurevich et al., 2004; Kuriyama et al., 2005), resembling the reduced or impaired vagal function in cardiac patients (Camm et al., 1996; Bauer et al., 2006). In contrast, because of degeneration of the preganglionic sympathetic neurons, it is hypothesized that the non-Gaussianity of HRV fails to markedly increase, such as that observed in cardiac patients (Kiyono et al., 2008; Hayano et al., 2011), in MSA patients. We test this hypothesis by comparing the results for MSA with those for CHF (Kiyono et al., 2008); the results were reanalyzed in the same methodological framework.

In the present study, we also studied ambulatory HRV in patients with Parkinson disease (PD) in which autonomic failure is commonly observed (Lipp et al., 2009). As the autonomic pathology of PD is different from that of MSA, being primarily postganglionic as evidenced by decreased uptake of adrenergic markers such as iodine-123 metaiodobenzylguanidine (Braune et al., 1998, 1999), the degree, and balance of sympathetic and vagal impairments could be different. Thus, it would be intriguing to examine if the lack of increased non-Gaussianity is still observed in PD.

**MATERIALS AND METHODS**

**STUDY PATIENTS**

Twelve MSA patients (six male and six female subjects; 61.9 ± 7.1, 54–76 years) and 10 patients with PD (two male and eight female subjects; 71.1 ± 6.0, 63–81 years) at the Department of Neurology of the University of Tokyo Hospital participated in this study (Tables 1 and 2, respectively). Diagnosis was made according to the UK Parkinson’s Disease Society Brain Bank Clinical Diagnostic Criteria (Hughes et al., 1992) and the second consensus statement on MSA diagnosis (Gilman et al., 2008). All patients were diagnosed with definite MSA or PD according to the published criteria.

**Table 1 | Clinical characteristics of multiple system atrophy (MSA) patients.**

| No | Age (years) | Sex | Clinical diagnosis | Symptoms at onset | Illness duration (years) | Ataxia | Parkinsonism | Autonomic failure |
|----|-------------|-----|-------------------|-------------------|--------------------------|--------|--------------|------------------|
| 1  | 60          | F   | MSA-C             | Instability of gait | 4                        | +++    | –            | +                |
| 2  | 64          | M   | MSA-P             | Gait disturbance   | 1                        | –      | ++           | +                |
| 3  | 61          | M   | MSA-C             | Dysautonomia       | 3                        | –      | ++           | +                |
| 4  | 57          | M   | MSA-C             | Orthostatic symptoms | 9                       | +      | –            | +                |
| 5  | 64          | F   | MSA-C             | Instability of gait | 3                        | ++     | +            | +                |
| 6  | 54          | M   | MSA-P             | Dysthria           | 4                        | –      | ++           | +                |
| 7  | 75          | F   | MSA-C             | Urinary urgency    | 4                        | ++     | –            | +                |
| 8  | 56          | F   | MSA-C             | Instability of gait | 2                        | ++     | –            | +                |
| 9  | 60          | M   | MSA-C             | Dysarthria, gait disturbance | 2            | ++     | +            | +                |
| 10 | 61          | F   | MSA-C             | Gait disturbance   | 4                        | ++     | ++           | +                |
| 11 | 55          | M   | MSA-C             | Instability of gait | 2                        | ++     | ++           | +                |
| 12 | 76          | F   | MSA-C             | Orthostatic symptoms | 2                       | +      | –            | +                |

MSA-C, MSA with predominant cerebellar ataxia; MSA-P, MSA with predominant Parkinsonism.

**Table 2 | Clinical characteristics of patients with Parkinson disease.**

| No | Age (years) | Sex | Clinical diagnosis | Symptoms at onset | Illness duration (years) | Drug | Hoehn–Yahr score |
|----|-------------|-----|-------------------|-------------------|--------------------------|------|-----------------|
| 1  | 68          | M   | PD                | Tremor            | 3                        | D, AC, DA | III |
| 2  | 63          | M   | PD                | Hand tremor       | 21                       | D, M, DA | IV  |
| 3  | 75          | F   | PD                | Hand tremor, gait disturbance | 11                   | D, DA, AM | IV  |
| 4  | 66          | F   | PD                | Gait disturbance, dysarthria | 13                   | D, DA | III |
| 5  | 75          | F   | PD                | Tremor, gait disturbance | 31                   | D, DA, AM, AC | IV |
| 6  | 81          | F   | PD                | Gait disturbance   | 5                        | D     | V   |
| 7  | 68          | F   | PD                | Tremor            | 8                        | D, AC, AM, DA, M | III |
| 8  | 65          | F   | PD                | Hand Tremor       | 8                        | D, DA | III |
| 9  | 72          | F   | PD                | Gait disturbance   | 5                        | D, DA, AC, AM | IV |
| 10 | 78          | F   | PD-D              | Gait disturbance, dysarthria | 2                       | D     | III |

PD, Parkinson disease; PD-D, PD with dementia; D, L-DOPA/carbidopa or benserazide; DA, dopamine agonists; A, anticholinergic; AM, amantadine; M, selegiline.
examine by neurologists, and all PD patients exhibited a response to L-DOPA without remarkable MRI findings. All MSA patients fulfilled the criteria for probable MSA (Gilman et al., 2008), and most of them took adrenergic stimulants for controlling severe orthostatic hypotension and anti-adrenergic or anti-muscarinic medications for their neurogenic bladder.

In addition, we studied 108 patients who were consecutively referred for evaluation or treatment of CHF (61 male and 47 female subjects; 66.1 ± 14.8, 21–92 years). Of these patients, 39 (36.1%) died within the follow-up period of 33 ± 17 months (range, 1–59 months). The medication status before discharge from the hospital was not significantly different between survivors and non-survivors. The clinical details of the CHF patients were reported previously (Kiyono et al., 2008).

The results were compared with data from age-matched healthy subjects; the details of which were reported elsewhere (Kiyono et al., 2004). All individuals within ±2 years of each patient’s age were selected from a pool of 122 healthy subjects.

MEASUREMENTS AND PROTOCOL
The original electrocardiogram (ECG) data were derived from 24-h Holter recordings. The ECG signals were digitized at 125 Hz and 12 bits and processed offline using a personal computer equipped with a dedicated software. All QRS complexes in each recording were detected and labeled automatically. The results of automatic analysis were reviewed, and any errors in R wave classification were corrected manually. Computer files were generated containing the duration of individual R–R intervals and morphology classifications of individual QRS complexes (normal, supraventricular, and ventricular premature complexes). The series of intervals between two consecutive R waves of sinus rhythm [normal-to-normal (NN) intervals] was analyzed. To avoid the adverse effects of any remaining errors in the detection of the R wave, large (>20%) consecutive R–R interval differences were thoroughly reviewed until all errors were corrected. In addition, when atrial or ventricular premature complexes were encountered, the corresponding R–R intervals were interpolated by the median of the two successive beat-to-beat intervals. We also confirmed that no sustained tachycardias were present in our HRV recordings. In this study, all HRV indices were obtained from the daytime (12:00–18:00) data.

ANALYSIS OF CONVENTIONAL HRV INDICES
The following HRV indices were calculated: mean NN intervals, standard deviation (SD) of all NN intervals (SDNN), SD of 5 min averaged NN intervals (SDANN), root mean square of successive difference of NN intervals (RMSSD), the variances corresponding to ultra-low frequency (ULF; 0–0.0033 Hz), very low frequency (VLF; 0.0033–0.04 Hz), low frequency (LF; 0.04–0.15 Hz), and high frequency (HF; 0.15–0.40 Hz) bands, and LF/HF ratio, all of which were proposed by the Task Force of the European Society of Cardiology and the North American Society of Pacing and Electrophysiology (Camm et al., 1996). The variances of these frequency components were transformed to natural logarithmic values (ln ms²).

In addition, we also computed the deceleration and acceleration capacity (DC and AC) based on the phase rectified signal averaging of NN intervals (Bauer et al., 2006), and the fractal scaling exponents, the short-term exponent α1 and the long-term α2, using detrended fluctuation analysis (DFA; Peng et al., 1995).

MULTISCALE PROBABILITY DENSITY FUNCTION ANALYSIS
Recent studies from our group have shown that human HRV exhibits the intermittent dynamics or temporal heterogeneity of variance leading to non-Gaussian probability density function (PDF) of heart rate increments (Kiyono et al., 2004), especially in cardiac patients within timescales corresponding to LF and VLF ranges (Kiyono et al., 2008; Hayano et al., 2011). As such a feature, called heteroscedasticity, cannot be captured by conventional HRV indices, we conducted multiscale PDF analysis to characterize intermittent large deviations and the resultant non-Gaussianity of HRV.

The procedure starts from interpolating observed series of NN intervals with a cubic spline function and resampling at an interval (Δt) of 250 ms (4 Hz), yielding interpolated time series b(t). Next after subtracting average interval b(avg), integrated time series B(t) are obtained by integrating b(t) over the entire length,

\[ B(t) = \frac{t}{\Delta t} \sum_{i=1}^{t/\Delta t} [b(i\Delta t) - b_{avg}]. \]

As in previous studies (Kiyono et al., 2004, 2008; Hayano et al., 2011), the local trend of B(t) is eliminated by a third-order polynomial fit to B(t) within moving windows of length 2s, where s is the scale of analysis. Thereafter, intermittent deviation \( \Delta_s B(t) \) is measured as the increment with a time lag s of the detrended time series. For instance, in a window from time \( T - s \) to \( T + s \), the increments are calculated as follows:

\[ \Delta_s B(t) = \left[ B(t + s/2) - f_{fit}(t + s/2) \right] - \left[ B(t - s/2) - f_{fit}(t - s/2) \right] \]

where \( T - s \leq t < T + s/2 \) and \( f_{fit}(t) \) is the polynomial representing the local trend of \( B(t) \). \( \Delta_s B(t) \) reflects an average degree of tachycardia if negative \( [b(t) < b_{avg}] \) or bradycardia if positive \( [b(t) > b_{avg}] \) over a moving window with length s (in seconds) after detrending. To quantitatively characterize the non-Gaussian property of \( \Delta_s B(t) \) at scale s, the standardized PDF (variance set to one) constructed from all \( \Delta_s B(t) \) values is approximated by the Caistaing’s model (Caistaing et al., 1990) with a single parameter \( \lambda_s \), which we refer to as the non-Gaussianity index. A greater \( \lambda_s \) indicates a fatter non-Gaussian tail and a sharper peak of PDF compared to the Gaussian distribution. On the other hand, if \( \lambda_s \) is close to zero, PDF is close to a Gaussian distribution. The parameter \( \lambda_s \) is estimated as follows:

\[ \lambda_s^2 = \frac{2}{q(q-2)} \left[ \ln \left( \sqrt{\pi} \left| \langle \Delta_s B \rangle \right|^q \right) - \ln \left( \frac{q+1}{2} \right) - \frac{q}{2} \ln 2 \right], \]

where \( q \neq 0 \) or 2, \( q > -1 \), and \( \langle |\Delta_s B| \rangle \) denotes the estimated value of the q-th order absolute moment of \( \Delta_s B \) (Kiyono et al., 2007).

In the present study, we calculated \( \lambda_s \) using the 0.25-th order moment (\( q = 0.25 \)) to emphasize the center part of PDF and
reduce the effects of large outliers such as those resulting from ectopic beats. This implies that our non-Gaussianity index with $q = 0.25$ more strongly characterizes peak PDF around the center of the observed non-Gaussian distribution, as opposed to higher-order moments, such as kurtosis based on the fourth moment, emphasizing heavy tails and extreme deviations. Based on our recent findings that increased $\lambda_s$ at scale $s = 25$ s is associated with increased cardiac mortality risk and that this predictive power is independent of clinical risk factors in CHF and AMI patients (Kiyono et al., 2008; Hayano et al., 2011), we evaluated the non-Gaussianity index $\lambda_{25s}$ at $s = 25$ s, which is at the edge of LF and VLF ranges.

An important feature of this multiscale PDF analysis is that if a time series has temporally homogeneous and finite variance, the increment PDF of the integrated series rapidly converges to a Gaussian distribution as the time-scale $s$ increases because of the well-known statistical law called the central limit theorem. On the other hand, if neither condition is fulfilled, slow convergence to a Gaussian distribution or a scale-dependent $\lambda_s$ and non-Gaussian fat tail can arise, suggestive of increased intermittency as observed in hydrodynamic turbulence (Castaing et al., 1990; Ghashghaie et al., 1996). Indeed, in the so-called multiplicative cascade model (Monin and Yaglom, 1975), one of the representative models describing intermittency of hydrodynamic turbulence and also used as a model of heart rate intermittency (Lin and Hughson, 2001), $\lambda_s$ is known to have scale dependence in the form of $\lambda_s^2 \sim \ln s$ (Kiyono et al., 2007; Figure 1). In the cascade model, multiscaling properties of the increments called structure functions also exist in the corresponding scales (Kiyono et al., 2007). To evaluate such a dynamic (cascade-like) aspect of intermittent fluctuations, we calculated the slope of $\lambda_s^2$ vs. $\ln s$ ($\lambda^2$-slope) in the range $20 < s < 200$ s (mainly covering LF and VLF ranges).

**STATISTICAL ANALYSIS**

The data are reported as the mean ± SD. One-way ANOVA was used to test for statistical differences across groups, and Tukey’s honestly significant difference test was used for pair-wise comparisons. For variables with skewed distributions, values were transformed to natural logarithms. The Kolmogorov–Smirnov test was used to assess differences in age distribution between groups. In addition, the bootstrap method (Efron and Tibshirani, 1993) was used to assess possible selection biases of age-matched control groups. Bootstrap samples having the same size as each of MSA and PD groups were generated by randomly drawing age-matched subjects with replacement from a pool of healthy subjects. $P < 0.05$ was considered significant.

**RESULTS**

Indices of autonomic function were derived from HRV recordings from MSA, PD, and CHF patients as well as from the three separate age-matched control groups (MSA controls, 63.6 ± 8.6 years vs. 65.8 ± 7.9 years).
MSA patients, 62.3 ± 7.4 years; PD controls, 68.5 ± 8.3 years vs. PD patients, 68.6 ± 7.9 years; CHF controls, 59.1 ± 16.0 years vs. CHF patients, 66.1 ± 14.8 years). The age distributions for the control groups were not significantly different from those for the patients’ groups. Mean duration since MSA diagnosis was 3.3 ± 2.1 years (range, 1–9 years; Table 1). Mean duration since PD diagnosis was 10.7 ± 9.1 years (range, 2–31 years), and the mean Hoehn and Yahr score was 3.6 ± 0.7 (range, 3–5; Table 2).

CONVENTIONAL HRV INDICES

Table 3 presents HRV indices derived from HRV recordings from MSA patients and age-matched healthy control subjects, together with the bootstrap estimators for the healthy controls. Compared with the control group, the MSA patients showed significantly decreased HRV as indicated by lower SDNN and SDANN, and RMSSD values, reduced power in all spectral bands (HF, LF, VLF, ULF), and lower DC and AC. Indices such as LF/HF and DFA $\alpha_1$ were also significantly decreased. Compared with the control group, the PD patients showed significant decreases only in LF and VLF power and significantly lower DC and AC (Table 4). LF/HF and DFA $\alpha_1$ were significantly decreased. As shown in Tables 3 and 4, these findings were largely supported also by comparing mean values for the patient groups with 95%-confidence intervals of the bootstrap estimators. Table 5 presents the HRV indices in CHF patients and age-matched healthy control subjects. Compared with the control group, both surviving and non-surviving CHF patients exhibited significantly decreased HRV as indicated by lower SDNN and SDANN, reduced power in LF, VLF, and ULF ranges, and lower

### Table 3 | Heart rate variability measures in patients with multiple system atrophy (MSA) and age-matched controls.

|                      | MSA (n = 12) | Age-matched control (n = 69) | P value | Bootstrap samples of age-matched control (n = 12) |
|----------------------|--------------|-----------------------------|---------|-------------------------------------------------|
| Mean NN, ms          | 766 ± 89     | 775 ± 110                   | 0.745   | 776 (723–832)                                   |
| SDNN, ms             | 59.7 ± 23.0  | 90.4 ± 28.6                 | <0.001  | 89.0 (78.4–104.2)                               |
| SDANN, ms            | 19.9 ± 6.5   | 47.5 ± 28.7                 | <0.001  | 48.9 (35.5–64.5)                                |
| RMSSD, ms            | 13.8 ± 4.4   | 22.3 ± 11.4                 | <0.001  | 21.8 (16.3–270)                                 |
| ln HF, ln ms$^2$     | 3.75 ± 0.90  | 4.97 ± 1.08                 | <0.001  | 4.93 (4.34–5.50)                                |
| ln LF, ln ms$^2$     | 4.02 ± 0.90  | 5.90 ± 0.97                 | <0.001  | 5.90 (5.39–6.36)                                |
| ln VLF, ln ms$^2$    | 5.92 ± 0.84  | 7.26 ± 0.81                 | <0.001  | 7.30 (6.89–7.72)                                |
| ln ULF, ln ms$^2$    | 7.78 ± 0.93  | 8.47 ± 0.64                 | 0.029   | 8.45 (8.14–8.79)                                |
| LF/HF ratio          | 1.69 ± 1.24  | 3.28 ± 2.49                 | 0.002   | 3.47 (2.25–4.87)                                |
| DC, ms               | 3.38 ± 0.98  | 6.23 ± 1.59                 | <0.001  | 6.82 (5.11–6.53)                                |
| AC, ms               | −3.38 ± 0.93 | −6.51 ± 1.77                | <0.001  | −6.13 (−6.94 to −5.28)                          |
| $\alpha_1$           | 0.86 ± 0.24  | 1.17 ± 0.15                 | <0.001  | 1.21 (1.08–1.33)                                |
| $\alpha_2$           | 1.23 ± 0.09  | 1.18 ± 0.04                 | 0.118   | 1.19 (1.15–1.23)                                |
| $\lambda_{255}$      | 0.46 ± 0.07  | 0.39 ± 0.07                 | 0.005   | 0.38 (0.35–0.43)                                |
| $\chi^2$-slope       | −0.05 ± 0.12 | −0.01 ± 0.08                | 0.309   | 0.00 (−0.04 to 0.04)                            |

Fifth column shows mean value and 95%-confidence interval based on 2000 bootstrap samples. $P < 0.05$.

### Table 4 | Heart rate variability measures in patients with Parkinson disease and age-matched controls.

|                      | Parkinson disease (n = 10) | Age-matched control (n = 60) | P value | Bootstrap samples of age-matched control (n = 10) |
|----------------------|---------------------------|-----------------------------|---------|-------------------------------------------------|
| Mean NN, ms          | 779 ± 118                 | 780 ± 112                   | 0.975   | 801 (717–885)                                   |
| SDNN, ms             | 70.4 ± 33.5               | 91.6 ± 29.6                 | 0.086   | 95.8 (76.1–118.7)                               |
| SDANN, ms            | 34.0 ± 26.7               | 46.6 ± 274                  | 0.191   | 44.4 (30.3–62.2)                                |
| RMSSD, ms            | 18.2 ± 11.6               | 23.1 ± 12.8                 | 0.240   | 26.3 (17.9–35.5)                                |
| ln HF, ln ms$^2$     | 4.21 ± 1.19               | 4.96 ± 1.13                 | 0.089   | 5.00 (4.30–5.71)                                |
| ln LF, ln ms$^2$     | 4.22 ± 1.31               | 5.70 ± 0.97                 | 0.006   | 5.69 (5.21–6.17)                                |
| ln VLF, ln ms$^2$    | 5.82 ± 1.20               | 7.17 ± 0.82                 | 0.006   | 7.13 (6.68–7.56)                                |
| ln ULF, ln ms$^2$    | 8.15 ± 0.77               | 8.55 ± 0.64                 | 0.148   | 8.66 (8.25–9.09)                                |
| LF/HF ratio          | 1.30 ± 0.13               | 2.82 ± 1.69                 | 0.003   | 2.48 (1.60–3.54)                                |
| DC, ms               | 3.93 ± 1.41               | 5.46 ± 1.60                 | 0.008   | 5.27 (4.45–6.11)                                |
| AC, ms               | −3.03 ± 1.57              | −5.80 ± 1.81                | 0.007   | −5.74 (−6.80 to −4.73)                          |
| $\alpha_1$           | 0.83 ± 0.26               | 1.12 ± 0.24                 | 0.011   | 1.07 (0.96–1.19)                                |
| $\alpha_2$           | 1.17 ± 0.06               | 1.19 ± 0.07                 | 0.456   | 1.18 (1.14–1.23)                                |
| $\lambda_{255}$      | 0.42 ± 0.09               | 0.40 ± 0.08                 | 0.574   | 0.41 (0.37–0.46)                                |
| $\chi^2$-slope       | −0.01 ± 0.11              | −0.02 ± 0.09                | 0.81    | −0.02 (−0.08 to 0.03)                           |

Fifth column shows mean value and 95%-confidence interval based on 2000 bootstrap samples. $P < 0.05$. 

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Table 5 | Heart rate variability indices in patients with congestive heart failure and age-matched controls.

|                  | CHF (NS; n = 39) | CHF (SV; n = 69) | Control (n = 90) | P value NS–SV | P value NS–C | P value SV–C |
|------------------|------------------|------------------|------------------|---------------|--------------|--------------|
| Mean NN, ms      | 758 ± 114        | 755 ± 140        | 782 ± 110        | 0.99          | 0.56         | 0.37         |
| SDNN, ms         | 59.1 ± 31.1      | 59.7 ± 39.6      | 93.4 ± 29.0      | 0.99          | <0.001       | <0.001       |
| SDANN, ms        | 33.0 ± 272       | 37.7 ± 33.6      | 51.6 ± 30.0      | 0.73          | 0.006        | 0.016        |
| RMSSD, ms        | 37.6 ± 40.0      | 40.0 ± 48.3      | 24.5 ± 12.8      | 0.94          | 0.13         | 0.017        |
| ln HF, ln ms2    | 5.43 ± 157       | 5.30 ± 167       | 5.13 ± 11.2      | 0.89          | 0.82         | 0.75         |
| ln LF, ln ms2    | 4.97 ± 180       | 4.85 ± 163       | 5.97 ± 101       | 0.90          | 0.001        | <0.001       |
| ln VLF, ln ms2   | 5.73 ± 136       | 6.04 ± 145       | 7.33 ± 83        | 0.40          | <0.001       | <0.001       |
| ln LF/HF         | 0.83 ± 0.79      | 0.93 ± 0.78      | 3.01 ± 2.31      | 0.95          | <0.001       | <0.001       |
| DC, ms           | 3.39 ± 160       | 3.84 ± 2.01      | 5.87 ± 1.70      | 0.43          | <0.001       | <0.001       |
| AC, ms           | −4.34 ± 2.29     | −4.69 ± 2.13     | −6.27 ± 1.96     | 0.68          | <0.001       | <0.001       |
| α1               | 0.79 ± 0.26      | 0.72 ± 0.24      | 1.17 ± 0.25      | 0.44          | <0.001       | <0.001       |
| α2               | 0.93 ± 0.16      | 1.00 ± 0.21      | 1.18 ± 0.08      | 0.048         | <0.001       | <0.001       |
| λ256             | 0.57 ± 0.18      | 0.48 ± 0.15      | 0.40 ± 0.08      | <0.001        | <0.001       | <0.001       |
| λ2-slope         | −0.21 ± 0.23     | −0.13 ± 0.18     | −0.02 ± 0.08     | 0.03          | <0.001       | <0.001       |

P< 0.05.

DC and AC. Indices such as LF/HF and DFA α1 and α2 were also significantly decreased.

In the MSA patients, the pattern of changes in conventional HRV indices was similar to that observed in the CHF patients. While the decreased HRV in both MSA and CHF patients might reflect reduced vagal heart rate control, decreases in LF/HF and DFA α1 were observed for both MSA, a disease with reported preganglionic sympathetic failure (Sone et al., 2005), and for CHF, a pathology associated with sympathetic overdrive (Packer, 1988; Ciarka et al., 2008). In contrast, no decreases in SDNN and HF power, indices of reduced HRV, were observed in the PD patients, which might reflect relatively intact vagal heart rate control. However, decreases in LF/HF and DFA α1 were also observed in PD, a disease with reported postganglionic sympathetic failure (Braune et al., 1998, 1999).

NON-GAUSSIAN AND INTERMITTENT PROPERTIES OF HRV

Figure 2 shows representative results of the multiscale PDF analysis for MSA, CHF, and PD patients (one patient from each group) and a healthy subject. As shown in Figures 2M–P, HRV data from the MSA and PD patients and the healthy subject yielded similar PDF curves at each scale. In contrast, recordings from the CHF patient yielded a PDF curve with a more tapered center and fatter tails at relatively smaller scales. This reflects intermittent large deviations or bursts observed at s = 20 s in CHF patients (Figure 2G), while this increased intermittency was not observed in the MSA and PD patients. In addition, as the scale s increases, deformation of PDFs toward a Gaussian distribution was clearly observed only in the CHF patient. The deformation process of the non-Gaussian PDF can be described by the relation between the non-Gaussianity index λ and scale s. As shown in Figure 3, the MSA and PD patient groups and the healthy subject groups showed nearly constant λ2 values across a wide range of scales s, resulting in an almost zero λ2-slope. In contrast, the CHF patient group, particularly non-survivors, was characterized by almost linear increases in λ2 as the log scale decreased from 200 to 20 s, similar to that observed for a cascade model of intermittent turbulence (Figure 1B). Consequently, the λ2-slope for the CHF patients was significantly more negative than that for the healthy controls. λ256 for the MSA patients was slightly but significantly higher than that for healthy controls (Table 3), although the level was much lower than that for CHF non-survivors (Table 5). λ256 for the PD patients failed to increase compared with that for healthy controls (Table 4). Both MSA and PD patients with sympathetic failure had λ2-slopes of almost zero, which were not significantly different from those of healthy controls (Tables 3 and 4). Only CHF patients with known sympathetic overdrive had significantly negative λ2-slopes (Table 5).

DISCUSSION

Long-term ambulatory HRV continues to attract clinical interest as a useful tool for risk stratification in AMI (Kleiger et al., 1987; Bigger et al., 1996; La Rovere et al., 1998; Schmidt et al., 1999; Huikuri et al., 2000; Bauer et al., 2006) and chronic heart failure (Ho et al., 1997; Nolan et al., 1998; Måkikallio et al., 2001). Patients at higher mortality risk frequently have higher heart rates with reduced and less complex (or monotonic) HRV, and most indices used to characterize such HRV dynamics primarily reflect reduced or impaired vagal function (Camm et al., 1996; Marine et al., 2002; Bauer et al., 2006). In contrast, few HRV indices are related to sympathetic function and their autonomic correlates and prognostic significance are still uncertain. For example, a decrease, but not the increase, in LF/HF, believed to reflect the sympathovagal balance (Pagani et al., 1986), is associated with increased mortality risk (Tsuij et al., 1994; Huikuri et al., 2000) in patients exhibiting sympathetic activation (Ciarka et al., 2008). Similarly, a decrease in DFA α1, known to be correlated with LF/HF and sensitive to changes in the sympathovagal balance (Huikuri et al., 2009), is associated with increased mortality risk (Huikuri et al., 2000; Måkikallio et al., 2001). The present study further demonstrated that both of these indices are also decreased in MSA, a neurodegenerative disorder associated with preganglionic sympathetic failure.
FIGURE 2 | Multiscale PDF characterization of heart rate variability. Illustrative examples of time series of NN intervals $b(t)$ (A–D), time series of $\Delta b_i(\sigma_s)$ (E–H), and standardized PDFs (in logarithmic scale) of $\Delta b_i(\sigma_s)$ for (from the top to bottom) $s = 20, 60, 180, 300$ s (M–P), where $\sigma_s$ denotes the SD of $\Delta b_i(\sigma_s)$. In solid lines, we superimpose the PDF approximated by Castaing’s model (Castaing et al., 1990). The panels on the leftmost side (A,E,I,M) are data for a 60-year-old male patient with multiple system atrophy (MSA). The panels on the left-of-center side (B,F,J,N) are data for a 68-year-old male patient with Parkinson disease (PD). The panels on the right-of-center side (C,G,K,O) are data for a 83-year-old female patient with congestive heart failure (CHF) who died 54 days after the measurement. The panels on the rightmost side (D,H,L,P) are data for a control subject (84-year-old male). For comparison, the dashed line denotes a Gaussian distribution. Note that, while the raw $b(t)$ for the patients looks much different from that for the control subject, the degrees of non-Gaussianity (i.e., the shapes of PDF) remain unaltered across scales for MSA and PD patients and the healthy control, except for those for the CHF patient at shorter scales (G,O).

(Sone et al., 2005), and in PD, which is often accompanied by postganglionic sympathetic failure (Braune et al., 1998, 1999).

As a marker potentially related to sympathetic cardiac overdrive, we have recently introduced increased non-Gaussianity of HRV within LF and VLF ranges in patients with CHF and AMI (Kiyono et al., 2008; Hayano et al., 2011), cardiopathologies known to be associated with sympathetic overdrive (Packer, 1988; Ciarreta et al., 2008). In the present study, we further demonstrated that a marked increase in intermittent and non-Gaussian HRV was not observed in MSA and PD patients with sympathetic failure. We still have not determined why $\lambda_{25s}$ for the MSA patients was slightly but significantly higher than that for healthy controls; this enhanced non-Gaussianity may be due to adrenergic stimulants administered to ameliorate severe orthostatic symptoms in the MSA patients. However, the scale-dependent increase in $\lambda_2$ with decreasing log scales mainly within the VLF range, leading to a markedly higher $\lambda_{25s}$ in the CHF patients (Table 5), was not observed in MSA. Therefore, we suggest that the systematically increased non-Gaussianity of HRV within LF and VLF ranges could be a hallmark of sympathetic cardiac overdrive and that indices such as $\lambda_{25s}$ and $\lambda^2$-slope could be used to measure the degree of sympathetic activation. Indeed, we recently observed decreased $\lambda_{25s}$ in the AMI patients taking (anti-sympathetic) $\beta$-blockers (Hayano et al., 2011).

Using concepts developed in statistical and non-linear physics, it has been demonstrated that the healthy human heart rate fluctuates in a complex manner even under resting conditions, exhibiting fractal long-range correlations (Peng et al., 1993; Yamamoto and Hughson, 1994) and multifractal properties (Ivanov et al., 1999; Amaral et al., 2001; Ching and Tsang, 2007). Based on these findings, Lin and Hughson (2001) proposed an analogy between heart rate dynamics and hydrodynamic turbulence because a
phenomenological model of hydrodynamic turbulence, called the multiplicative cascade model (Monin and Yaglom, 1975), can also have multifractal properties. Using multiscale PDF analysis, we later demonstrated that the healthy human HRV does not show slow and gradual convergence to a Gaussian distribution (Kiyono et al., 2004; Figure 3), an important requirement of the multiplicative cascade model (Figure 1B). In contrast, the present study and previous work (Kiyono et al., 2008) suggest that HRV within LF and VLF ranges of CHF patients, especially non-survivors, is more compatible with the multiplicative cascade model.

The multiplicative (log-normal) cascade model used to generate fluctuations with intermittent bursts such as those shown in Figure 1A is given by

\[ x_i = \xi_i \exp \left[ \sum_{j=1}^{m} \omega^{(j)} \left( \left\lfloor \frac{j-1}{2^{m-j}} \right\rfloor \right) \right], \]

where \( \xi_i \) is Gaussian white noise with zero mean, \( \omega^{(j)}(k) \) are independent Gaussian random variables with zero mean and constant variance, and \( \lfloor \cdot \rfloor \) is the floor function (Kiyono et al., 2007). The \( m \) is the total number of cascade steps, yielding the total number of data points \( 2^m \) (\( i = 1, \ldots, 2^m \)). An essential part of the model is that \( \xi_i \) is modulated by multiplication of random (log-normal) weights \( \exp[\omega^{(j)}(k)] \) (\( k = 0, 1, \ldots, 2^m-1 \)) at the \( j \)-th step every \( 2^{m-j} \) subintervals; therefore, large fluctuations are observed only when the momentary weights for (many) different steps with varying timescales are simultaneously large (refer to Figure 5 of Kiyono et al., 2007). Using multiscale PDF analysis, Kiyono et al. (2007) further showed that this model exhibits the scale dependence of a non-Gaussianity index in the form of \( \lambda_2^s \sim \ln s \) (Figure 1B).

The fact that heart rate dynamics of CHF patients with sympathetic activation exhibit a non-Gaussianity index which decays with scales within LF and VLF ranges suggests a sympathetic origin for HRV intermittency. In these scales (20–200 s), heart rate dynamics reflect cardiovascular regulation by neural, humoral, and thermal influences (Kitney and Rompelman, 1980). These subsystems are considered to be compensatory; therefore, it is likely that only simultaneous failure of all these subsystems operating at multiple timescales, compatible with the reciprocal of cascade steps \( \left\lfloor \frac{j}{2^{m-j}} \right\rfloor \) in the above example, could result in sympathetic overdrive, leading to large and intermittent heart rate deviations. We propose that such a multiplicative picture would provide a deeper physiological understanding of the nature of sympathetic function. In addition, it would provide a reason why methods requiring stationary, not intermittent, dynamics have not been successful in finding the sympathetic correlates of ambulatory HRV.

In the present study, we focused on daytime HRV for the following reasons. First, as reported in our previous study (Kiyono et al., 2005), there are large differences in non-Gaussianity and its scale dependence between day and night in healthy humans, presumably because of the difference in the sympathovagal balance. Second,
disorders of sleep and sleep breathing are common in MSA (Colosimo, 2011); therefore, incorporating nighttime data would inevitably introduce additional complexity. Third, one of our goals is to assess sympathetic activity, which is predominant during the day. Note that this shift from 24-h HRV to daytime HRV does not change our previous finding of the increased non-Gaussianity of low frequency HRV in CHF patients than in healthy controls (Kiyono et al., 2008).

In agreement with previous studies (Gurevich et al., 2004; Kuriyama et al., 2005), our MSA patients showed significantly decreased HRV, as evidenced by lower SDNN and HF power. This decrease is probably related to the known abnormalities in central vagal (Benarroch et al., 2006) and sympathetic function in these patients (Sone et al., 2005). On the other hand, changes in SDNN and HF power were not significant in PD patients, implying relatively intact vagal heart rate control despite the impaired peripheral, cardiac sympathetic function (Braune et al., 1999, 1999). Thus, analyses of ambulatory HRV may facilitate discriminative diagnosis between MSA and PD, particularly the difficult distinction between early stage PD and MSA with predominant Parkinsonian symptoms (Lipp et al., 2009).

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