Annual seasonal influenza vaccination (SIV) is recommended for people with diabetes, but vaccine coverage remains low. We estimated the probabilities of stopping or starting SIV, their correlates, and the expected time spent in the vaccinated state over 10 seasons for different patient profiles. We set up a retrospective cohort study of patients with diabetes in 2006 (n = 16,026), identified in a representative sample of beneficiaries of the French National Health Insurance Fund. We followed them up over 10 seasons (2005/06–2015/16). We used a Markov model to estimate transition probabilities and a proportional hazards model to study covariates. Between two consecutive seasons, the probabilities of starting (0.17) or stopping (0.09) SIV were lower than those of remaining vaccinated (0.91) or unvaccinated (0.83). Men, older patients, those with type 1 diabetes, treated diabetes or more comorbidities, frequent contacts with doctors, and with any hospital stay for diabetes or influenza during the last year were more likely to start and/or less likely to stop SIV. The mean expected number of seasons with SIV uptake over 10 seasons (range: 2.6–7.9) was lowest for women <65 years with untreated diabetes and highest for men ≥65 years with type 1 diabetes. Contacts with doctors and some clinical events may play a key role in SIV adoption. Healthcare workers have a crucial role in reducing missed opportunities for SIV. The existence of empirical patient profiles with different patterns of SIV uptake should encourage their use of tailored educational approaches about SIV to address patients’ hesitancy.

### Introduction

Because people with diabetes are at increased risk of severe complications linked to seasonal influenza, the World Health Organization (WHO) and many national guidelines recommend they receive annual seasonal influenza vaccination (SIV). The SIV rate in this population is nonetheless far below WHO’s target of 75% in most Western countries, including France (about 30% among those <65 years and 60% among those ≥65 years). Vaccine hesitancy among people with diabetes may contribute significantly to suboptimal vaccination coverage, as observed in other population groups, such as parents. The complexity of the vaccination pathway in France – despite efforts to simplify it – may also contribute to this situation.

SIV must be repeated annually because immunity from vaccination wanes over time and because of the constant evolution of influenza viruses. Each season, the decision to be vaccinated or not is ultimately the patient’s choice. Quantifying behavior changes (starting or stopping SIV) and understanding their determinants are both key to improving public health programs and patient education. Nonetheless, few longitudinal studies have explored the course of SIV behaviors over several consecutive years to obtain chronological information about the clinical events preceding and explaining SIV behavior changes. Multistate Markov (MSM) models, because they allow the analysis of transitions from one health state to another, are particularly appropriate for analyzing sequences of health events. These models make it possible to estimate transition probabilities, namely, the probability of remaining in the same state or of moving to another, and to identify factors affecting these transitions. We used this approach to understand the dynamics of SIV behavior changes and identify clinical profiles predictive of this behavior over 10 years.

Our objectives were: 1) to estimate SIV transition probabilities (see above); 2) to quantify associations between these transitions and sociodemographic, clinical, and healthcare utilization characteristics; and 3) to assess the total time spent in the vaccinated state over the 10-season study period according to patients’ characteristics. This third objective seeks to determine easily identifiable demographic-clinical profiles to help clinicians tailor SIV promotion strategies to their patients.

### Materials and methods

**Study design and data source**

We conducted the study in the Permanent Sample of Beneficiaries (Echantillon Généraliste des Bénéficiaires, a,b,c, S. Cortaredona d,e, L. Fressard a,b,c, F. Galtier e, and P. Verger a,b,c,d

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- iHU-Méditerranée Infection, Marseille, France; 
- ORS PACA, Observatoire Régional de la Santé Provence-Alpes-Côte d’Azur, Marseille, France; 
- INSERM, C-FRIN, Innovative Clinical Research Network in Vaccinology (I-Reivac), GH Cochin Broca Hôtel Dieu, Paris, France; 
- CIC 1411, CHU Montpellier, Hôpital Saint Eloi, Montpellier, France

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EGB). The EGB, set up in 2005, is a permanent, representative, and open national random sample of 1/97th of the people affiliated with one of the three major national health insurance funds in France (salaried workers, agricultural workers and farmers, and self-employed workers); these cover nearly 95% of the French population. In August 2017 (time of extraction), the database included 804,089 beneficiaries. For this study, we extracted data for salaried workers (including those who are retired) and their covered family members (i.e., children and non-working spouses), who are covered by the French National Health Insurance Fund (NHIF) and account for about 86% of the French population. Accessible data were age, sex, long-term illness (LTI) status, and reimbursement claims – for consultations with private healthcare professionals, medical procedures, and drugs purchased in the community (classified by Anatomical Therapeutic Chemical (ATC) codes). At their doctors’ requests, the NHIF grants LTI status to beneficiaries with long-term and costly diseases; this status exempts them from copayments for any medical care associated with that disease, regardless of their income level.

Data regarding the diagnoses associated with admissions since 2006 to a French public or private hospital, excluding military and psychiatric hospitals, are also available for all individuals in the EGB.

The study analyzed reimbursement data from the National System of Health Data (Système National des Données de Santé), in accordance with the General Conditions of Use of the Portal and the Data (Conditions Générales d’Utilisation du Portail et des Données) (version 3.0). Because the study was performed in accordance with Article L1461-1 (paragraph III.6) of the French Public Health Code (Code de la santé publique) for public health purposes with fully anonymized data, there were no further requirements for ethical approval, consent to participate, or data protection agency approval.

**Study population**

The study population comprised patients identified with diabetes (all types) in 2006, residing in mainland France, and with at least two years of follow-up over the study period. To identify them, we used an NHIF algorithm based on LTI status, hospitalization diagnoses and reimbursement claims for antidiabetic drugs or hemoglobin A1c (HbA1c) assays (Supplementary Table 1). Those who died or withdrew from the NHIF during the follow-up period were censored at the start of the year of the event.

**Outcome: SIV uptake**

For each individual and each SIV season n/n + 1 (temporal statistical units in our analyses), we constructed a binary variable "SIV uptake" (yes, no), based on SIV deliveries (ATC codes J07BB02, J07BB03, J07CA, excluding A/H1N1 pandemic vaccines, Supplementary Table 2) recorded between September 1 of year n and March 31 of year n + 1. Each SIV delivered in a community pharmacy is recorded in the NHIF database. SIV dispensing was considered a proxy for SIV uptake.

**Clinical characteristics of cohort members**

Cohort members without diabetes in 2005 according to the algorithm mentioned above were considered "patients with newly identified diabetes". Furthermore, we categorized each patient, each year, according to the characteristics and treatment of her/his diabetes, from data on LTI status and reimbursement claims. Specifically, the reimbursements were those for antidiabetic drugs recorded during the six months before the start of the season under consideration (n/n + 1): type 1 diabetes (LTI status for type 1 diabetes: E10 according to the 10th International Classification of Diseases, ICD-10, and/or treated by insulin at inclusion); other types not treated by antidiabetic drugs or insulin; other types treated by antidiabetic drugs and/or insulin.

To measure comorbidities, we calculated an individual chronic condition score (ICC) for each year of follow-up, based on drug deliveries (excluding antidiabetic drugs). This methodology has previously been published and used in several articles.

We also extracted, for each cohort member, the number of hospital stays each year (between September 1 of year n, and August 31 of year n + 1), for diabetes, diabetes complications, influenza, and influenza complications, and the numbers of visits – separately – with general practitioners (GPs), pulmonologists, endocrinologists, and cardiologists. GPs usually manage most patients with type 2 diabetes and refer patients to specialists when they face difficulties or complications. We also extracted information about patients’ changes of GP during follow-up.

**Statistical analysis**

For objective 1, we used an MSM model and estimated, based on transition intensity (i.e., the instantaneous risks of moving from one state to another), the transition probability \( P_{r,s}(t) \): namely, the probability of remaining in the same state \( r \) or moving from state \( r \) to another state \( s \) within a given period of time \( t \) (here, a 1-year period: from season n-1/n to season n/n + 1). Each season, we considered two states (vaccinated against seasonal influenza, or not), which generates four kinds of transitions (Figure 1): remaining vaccinated, remaining unvaccinated, starting vaccination, and stopping vaccination. Given the Markov property of our MSM model, the future SIV state is independent of the past state, given the present state and the covariate values. In other words, the probability of SIV the next season depends only on the current state (SIV or no SIV) and the covariate values.

As the A(H1N1) pandemic season (2009/10) had a major impact on SIV coverage rates, we could not use a standard MSM model that would assume constant transition intensities over time. We implemented a time-inhomogeneous model with piecewise-constant intensities that makes it possible to take into account specific points in time when an identifiable event (here the 2009 pandemic) may imply a change in intensities of transition for all subjects.

For objective 2, we used a proportional hazards model to model transition intensities as a function of constant (sex, age at inclusion) and time-varying covariates (type of diabetes, ICC...
score, hospital stays for diabetes/influenza, and healthcare utilization):

\[
q_n(z(t)) = q_n^{(0)} \exp(\beta^T z(t))
\]

where \(q_n\) is the transition intensity from state \(r\) to state \(s\); \(z(t)\) is the vector of covariates at time \(t\), and \(\beta\) is the vector of regression coefficients, which can be interpreted as a hazard ratio (HR). This model makes it possible to test the effect of covariates on transition intensities, even if some patients have multiple transitions from one state to another.\(^{16}\)

For objective 3, we used the transition intensities and covariate coefficient estimates to calculate the total time spent in the “SIV uptake” state over 10 SIV seasons. Given the initial state (SIV/no SIV uptake) and a set of three covariates (age, gender, and type of diabetes and its treatment), we were able to estimate the expected (i.e., predicted by the MSM model) number of seasons with SIV uptake over the next 10 seasonal campaigns for different patient profiles. We selected these three covariates because they were strongly associated with SIV uptake and are easily available to physicians in routine practice.

The goodness of fit of the model was assessed by the likelihood-ratio (LR) test and a graphical comparison of the predicted and observed prevalence rates of each state at each seasonal campaign. As most patients were lost during follow-up because of death, we also tested an empty (without covariates) three state model with “death” as an absorbing state to take the probability of dying during follow-up into account. This model yielded to similar results (results available upon request).

Descriptive statistical analyses were performed with SAS, version 9.4 (SAS Institute Inc., Cary, NC, USA). The MSM was fitted with R 3.5.0 software, and the MSM package version 1.6.6 developed by Jackson.\(^{19}\)

**Results**

**Study population characteristics (Table 1)**

Of the 17,259 patients identified with diabetes in 2006, we included in the cohort 16,206 patients (94%) with at least two years of follow-up over the study period. Mean age at inclusion was 65 ± 14 years, and 46% were women. Fewer than 10% were patients with diabetes newly identified at inclusion, and 10% had type 1 diabetes. The average duration of follow-up was 8.6 ± 2.4 years (see details in Supplementary Table 3). Over the follow-up period, 27% of the initial cohort died and 4% were lost to follow-up. SIV uptake rates were estimated at 51% in 2006/2007 and 57% in 2015/2016.

**Objective 1: transition probability estimates**

Over the study period, 75% of the patients had stable SIV behaviors (i.e., never vaccinated, always vaccinated, or had only one SIV transition) (Figure 2). Between two consecutive seasons, the probability of starting SIV (0.17, 95% confidence interval – CI – 0.17, 0.18) was significantly higher than the probability of stopping it (0.09, 95% CI 0.09, 0.10). The probability of remaining vaccinated (0.91, 95% CI 0.90, 0.91) was significantly higher than the probability of remaining nonvaccinated (0.83, 95% CI 0.82, 0.83). The goodness of fit of our model was satisfactory (Supplementary Figure 1).

**Table 1. Study cohort characteristics during the first and last season n/n + 1 of follow-up (EGB, France, 2006/07–2015/16, n = 16,206).**

| Description                                      | 2006/07 (n = 16,206) | 2015/16 (n = 11,277) |
|--------------------------------------------------|----------------------|----------------------|
| Average duration of follow-up in years           | 8.6 (2.4)            |                      |
| (min = 2; max = 10) – mean (SD)                   |                      |                      |
| **Sociodemographic characteristics**              |                      |                      |
| Age (years) on December 31 of year n – mean (SD) | 64.5 (13.5)          | 70.3 (12.7)          |
| Women                                            | 46.2                 | 47.1                 |
| **Clinical characteristics**                      |                      |                      |
| Patients with newly identified diabetes\(^{4}\)  | 8.9                  | NA                   |
| Type and treatment of diabetes                    |                      |                      |
| Type 1 diabetes\(^{6}\)                         | 9.5                  | 9.8                  |
| Other type diabetes treated by insulin and/or anti-diabetic drugs | 76.1                  | 81.8                |
| Other type diabetes without any treatment         | 14.4                 | 8.5                  |
| Weighted individual chronic condition score\(^{6}\) – mean (SD) | 0.8 (0.5)            | 0.9 (0.5)            |
| Hospitalized for diabetes or its complications between September 1 of year n-1 and August 31 of year n | 6.2                  | 4.3                  |
| Hospitalized for influenza or its complications between September 1 of year n-1 and August 31 of year n | 0.6                  | 1.3                  |
| **Healthcare use**                                |                      |                      |
| Number of consultations between September 1 of year n-1 and August 31 of year n, with – mean (SD) | 8.6 (7.1)           | 7.8 (6.3)           |
| General practitioner                              | 0.0 (0.3)            | 0.1 (0.6)            |
| Pulmonologist                                     | 0.3 (1.2)            | 0.4 (1.2)            |
| Endocrinologist                                   | 0.5 (1.4)            | 0.5 (1.4)            |
| Cardiologist                                      |                      |                      |
| Change of general practitioner between             | 4.0                  | 10.2                 |
| September 1 of year n-1 and August 31 of year n   |                      |                      |
| SIV uptake rate                                   | 50.8                 | 57.1                 |

Abbreviations: NA, not applicable; SD, standard deviation; SIV, seasonal influenza vaccination.\(^{4}\)

Among all patients included in the cohort, 4334 (26.7%) died during follow-up and 595 were lost to follow-up (3.7%).\(^{1}\)

**Additional notes:**

\(^{1}\)Patients with newly identified diabetes were those not detected as having had diabetes the year before inclusion.

\(^{4}\)People with type 1 diabetes were those with long-term illness status for type 1 diabetes (E10 according to the ICD-10) and treated by insulin at inclusion.

\(^{6}\)The individual chronic condition score (ICC) was calculated as a weighted sum of 21 chronic conditions. Weights account for the severity of each condition in the score calculation (ICC range in study cohort: min = 0; max = 3.7).
Objective 2: factors associated with starting or stopping SIV (Table 2)

Compared with patients who remained nonvaccinated, the probability of starting SIV was higher for men and increased with age. It was also higher for patients with newly identified diabetes in 2006, type 1 diabetes, more comorbidities, or hospitalized for diabetes or influenza since the previous SIV season. Patients with frequent contacts with GPs or at least one contact with a pulmonologist and those who had changed GP since the previous season were also more likely to start SIV.

Compared with patients who remained vaccinated, the probability of stopping SIV was lower for men and decreased with age. It was also lower for patients with type 1 diabetes or another type of diabetes treated by insulin and/or antidiabetic drugs. Changing GP increased the probability of stopping, as well as starting, SIV, and contacts with endocrinologists and cardiologists reduced the probability of stopping SIV.

Objective 3: expected time spent in the “SIV uptake” state, according to patient profiles

For patients not vaccinated at the beginning of the period, the group with the lowest expected number of seasons with SIV uptake over the 10-season period (Figure 3a) (mean = 2.6 seasons) was that of women younger than 65 years with untreated diabetes, and the group with the most vaccinated seasons (mean = 5.7) was that of men aged 65 years or older with type 1 diabetes.

For patients vaccinated at the beginning of the period, the expected number of seasons with SIV uptake was higher for all profiles, but their ranking remained similar (Figure 3b).

Discussion

Among this cohort of patients with diabetes followed up over 10 seasons, the probability of remaining vaccinated or nonvaccinated was far higher than the probability of moving from one state to the other. The probability that a nonvaccinated patient would start SIV during the next season was higher (0.17) than the probability of a vaccinated patient stopping (0.09). Men, older patients, those with type 1 diabetes, treated diabetes or more comorbidities, frequent contacts with GPs or pulmonologists, or with any hospital stay for diabetes or influenza during the previous year were all more likely to start and/or less likely to stop SIV. Changing GP was associated with starting SIV among those non-vaccinated, and stopping it among those who had previously been vaccinated. The mean expected number of seasons with SIV uptake depended on initial SIV state and patients’ demographic and clinical profiles.

Our results demonstrate, consistently with previous findings, that people with diabetes have mainly stable SIV behaviors. The finding that the probability of remaining vaccinated (0.91) was significantly higher than that of remaining unvaccinated (0.83) is somewhat encouraging, even though the order of magnitude of these probabilities did not differ notably. Nonetheless, the probabilities of behavioral shifts were non-negligible in each group and higher for starting to vaccinate than for stopping it. Our results (Table 2) confirm that both modifiable (medical care) and nonmodifiable factors (aging and clinical deterioration) were associated with increased SIV uptake rates. This finding is both intuitive and fortunate and helps to lighten, in part, the dark picture of low SIV coverage in patients with diabetes: low coverage involves, above all, patients with the least serious forms of disease. Our results suggest that any clinical encounter with patients with diabetes, both vaccinated and nonvaccinated, should be taken as a new opportunity to make them aware of the importance of SIV, especially when their diabetes is becoming more severe. Our results also suggest that this is already the case, to some extent, as we found a higher probability of starting SIV among patients hospitalized for either influenza (+35%) or diabetes (+12%).


These results and their clinical implications are important for doctors. First, the probability of starting SIV rose with contacts with GPs and especially with pulmonologists in the last year. These findings confirm the key role of doctors in promoting SIV among at-risk patients. Previous studies have shown that patients trust their physicians strongly and that these professionals can play an important role in improving patients’ knowledge and decision-making process about SIV. The absence of any similar effect for endocrinologists and cardiologists, however, suggests that these specialists might be missing some opportunities to promote SIV. In particular, it is surprising that medical contacts with endocrinologists are not associated with an increase in SIV uptake, as endocrinologists work with patients on behavioral changes in their daily practice, particularly concerning diet or physical activity. Endocrinologists and cardiologists might have a lower perception of the severity of influenza than pulmonologists because the association between influenza-induced complications of diabetes or cardiovascular diseases might appear less frequent, direct or severe than in patients with chronic respiratory diseases. These specialists may also have the perception that patients will receive SIV elsewhere, a common explanation for missed vaccination among some specialists. Increasing awareness among endocrinologists and cardiologists of the benefits of SIV for people with diabetes (e.g., through guidelines and initial and continuing medical education) is an important goal.

Second, doctors must be aware that some specific events (e.g., intensification of treatment and hospitalization for diabetes or influenza) may present favorable opportunities for encouraging SIV uptake. These events may increase patients’ perception of their own vulnerability toward influenza or of the severity of the disease – both strong facilitators of the adoption of SIV. Moreover, they provide a good practical opportunity to meet caregivers involved in therapeutic education programs, provided that these include education about SIV.

Third, our results also suggest that the first visits of new patients may be a critical period for doctors to review their prevention needs (as shown for medications and promote SIV. We also found, however, that changing GPs may be correlated with stopping SIV as well. Lack of additional data on GPs’ characteristics makes this finding difficult to interpret. Changing GPs might reflect a change in a patient’s commitment to health or attitudes toward healthcare, and doctors should be particularly attentive to these new patients.

Fourth, our results highlight the existence of profiles of patients with contrasting propensities for regular vaccination. For example, we estimated that a woman with diabetes younger than 65 years and with no pharmacological treatment for this disease will be vaccinated only about 3 times over the next 10 seasons (vs around 7 times for a woman 65 years or older with type 1 diabetes). While clinicians should systematically assess whether or not their patients have been vaccinated against the flu, they should also bear these profiles in mind to help them to identify those at high risk of irregular or no SIV uptake. This would help them to anticipate the need to spend more effort and time to focus on them, to understand their motives for potential hesitancy/reluctance, and to propose tailored approaches to address these motives and encourage SIV, as recommended for addressing vaccine hesitancy.

These results have also several implications for policymakers. According to the Community Guide’s categories of interventions to increase appropriate vaccination, strategies to improve SIV rates in target populations should rely on: increasing community demand for SIV (e.g., media campaigns, patient reminders); enhancing access to SIV (e.g., in France, free vaccine, vaccination by nurses since 2008, and by community pharmacists – nationwide implementation starting in 2019); and using provider- or system-based interventions (e.g., pay-for-performance for GPs, where performances includes SIV indicators). Applying these recommendations at the national level has remained insufficient in many countries, and SIV rates have remained far lower than the targets of both WHO and individual nations. Further efforts are required.
One method, implied by our results, may involve training to enhance the educational skills of physicians and other healthcare workers’ educational skills; it should also provide them with guidance and tools for adapting SIV educational efforts to different patient profiles. No “one size fits all” intervention is likely to exist, and tailored interventions adapted to patients’ characteristics, attitudes, and perceptions may be more effective in fostering SIV.

For example, providing communication training to healthcare workers, together with tailored information sheets and patient education websites, should be evaluated for increasing SIV uptake, as it has been shown to be an effective strategy for increasing uptake for HPV vaccine. As healthcare workers are also prone to vaccine hesitancy, and SIV coverage remains suboptimal among them in many countries, health authorities should be evaluated for increasing SIV uptake, as it has been shown to be an effective strategy for increasing uptake for HPV vaccine. As healthcare workers are also prone to vaccine hesitancy, and SIV coverage remains suboptimal among them in many countries, health authorities should be evaluated for increasing SIV uptake, as it has been shown to be an effective strategy for increasing uptake for HPV vaccine. As healthcare workers are also prone to vaccine hesitancy, and SIV coverage remains suboptimal among them in many countries, health authorities should be evaluated for increasing SIV uptake, as it has been shown to be an effective strategy for increasing uptake for HPV vaccine.

The strengths of this study include its large sample size, and the use of a combination of different criteria to identify patients with diabetes with various clinical features (Supplementary Table 1). Another strength of our study is its 10-year follow-up period, which is, with our previous article based on the same cohort, the longest of the studies of SIV behavior over several consecutive seasons. From a statistical standpoint, our study is the first to use an MSM model to estimate transition probabilities for SIV behaviors. Moreover, although several previous studies have examined the factors associated with being vaccinated (or not), our approach is the first one to explore associations between shifts in SIV behaviors and several demographic and clinical characteristics. This is essential for identifying opportunities (critical periods and events) to promote SIV. Our model provided a good prediction of SIV rates over the 10-year follow-up period (Supplementary Figure 1), confirming that the Markovian approach is well adapted to modeling SIV state transitions. Moreover, our time-inhomogeneous model with piecewise-constant intensities allowed us to adjust for some variations of the probabilities of SIV uptake during the follow-up period, especially those, now well documented, induced by the 2009 A(H1N1) campaign in France.

We also acknowledge some limitations. First, vaccinations that took place during occupational medicine visits or at vaccination centers or some nursing homes that buy vaccines for their residents (fewer than 20% of all nursing homes) are not recorded in the French NHIF databases. However, these limitations are unlikely to affect our results substantially as the vast majority of vaccinations in France are administered by private healthcare workers and are thus recorded in these databases. Second, as SIV behavior may vary by social characteristics, our results cannot be extrapolated to population categories (e.g., farmers, tradespeople, self-employed individuals) that could not be included in this study because they were first included in the EGB only in 2011. Nonetheless, salaried workers (including those who are retired, and their covered spouses and children) account for 86% of the French population. Third, our study is observational and retrospective and thus cannot draw definitive conclusions about causality. Notwithstanding this
limitation, our statistical approach is respectful of the chronological succession of events and therefore more robust because it also limits confounding by reverse causality. That is, SIV may be associated with a significant decrease in risk for hospital admission due to stroke, heart failure, and influenza or pneumonia, and an approach that fails to take the order of events into account might easily confound the association between hospitalization for diabetes or its complications and the probability of SIV uptake. Finally, several socioeconomic characteristics (e.g., educational level) and cognitive factors (e.g., beliefs about SIV efficacy and safety) may be associated with changes in SIV behaviors but could not be included in the analysis because they are not recorded in the NHIF databases.

While several trials have assessed the effectiveness of interventions to increase SIV rates (especially among older people), further research is needed to determine how best to target these interventions to specific populations, depending on patient characteristics. Further research is also needed to assess the effectiveness of interventions on SIV vaccination that have showed promising results for other vaccines and population groups (e.g., motivational interviewing and web-based tailored educational tools).

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Disclosure of potential conflicts of interest

The authors declare that they have no competing interests.

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