Association of noninvasive respiratory support with mortality and intubation rates in acute respiratory failure: a systematic review and network meta-analysis

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Abstract

Background: Noninvasive respiratory support devices may reduce the tracheal intubation rate compared with conventional oxygen therapy (COT). To date, few studies have compared high-flow nasal cannula (HFNC) use with noninvasive positive-pressure ventilation (NPPV). We conducted a network meta-analysis to compare the effectiveness of three respiratory support devices in patients with acute respiratory failure.

Methods: The Cochrane Central Register of Controlled Trials, MEDLINE, EMBASE, and Ichushi databases were searched. Studies including adults aged ≥ 16 years with acute hypoxic respiratory failure and randomized-controlled trials that compared two different oxygenation devices (COT, NPPV, or HFNC) before tracheal intubation were included. A frequentist-based approach with a multivariate random-effects meta-analysis was used. The network meta-analysis was performed using the GRADE Working Group approach. The outcomes were short-term mortality and intubation rate.

Results: Among 5507 records, 27 studies (4618 patients) were included. The main cause of acute hypoxic respiratory failure was pneumonia. Compared with COT, NPPV and HFNC use tended to reduce mortality (relative risk, 0.88 and 0.93, respectively; 95% confidence intervals, 0.76–1.01 and 0.80–1.08, respectively; both low certainty) and lower the risk of endotracheal intubation (0.81 and 0.78; 0.72–0.91 and 0.68–0.89, respectively; both low certainty); however, short-term mortality or intubation rates did not differ (0.94 and 1.04, respectively; 0.78–1.15 and 0.88–1.22, respectively; both low certainty) between NPPV and HFNC use.

Conclusion: NPPV and HFNC use are associated with a decreased risk of endotracheal intubation; however, there are no significant differences in short-term mortality.

Trial registration: PROSPERO (registration number: CRD42020139105, 01/21/2020)

Keywords: Acute hypoxic respiratory failure, Conventional oxygen therapy, Noninvasive ventilation, High-flow nasal cannula, Systematic review, Meta-analysis, Network meta-analysis

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Background
Acute respiratory failure (ARF) is prevalent among critically ill patients and is a common cause of intensive care unit (ICU) mortality [1, 2]. Approximately 60% of patients with ARF require invasive mechanical ventilation (IMV) [3], which is associated with adverse events, including ventilator-induced lung injury (VILI) and ventilator-associated pneumonia (VAP) [4, 5]. Patients with ARF on IMV have high hospital mortality rates of up to 30% [4]. Initial respiratory support, including conventional oxygen therapy (COT; e.g., nasal cannulas and facemasks), noninvasive positive-pressure ventilation (NPPV), and high-flow nasal cannula (HFNC) use, are important treatments to prevent tracheal intubation and reduce mortality among patients with hypoxic respiratory failure [6–10].

There is a widespread application of NPPV in patients with ARF before tracheal intubation and IMV [6–8] which decreases the need for IMV rather than the use of COT [11, 12]. NPPV potentially increases the risk of complications, including aspiration pneumonia, facial skin breakdown, eye irritation, interface intolerance, and patient discomfort from the inability to communicate or eat during therapy [13, 14], which limits NPPV application in the clinical setting. HFNC can deliver high-concentration humidified oxygen via nasal cannulas without NPPV-related complications and is increasingly used in critically ill adult patients despite contradictory results from several clinical trials [9, 10]. However, there is a paucity of evidence on pre-IMV HFNC use in patients with ARF.

Systematic reviews and meta-analyses that compared two of the three respiratory support devices (COT, NPPV, and HFNC) [15–22] showed that HFNC use reduced the tracheal intubation rate compared with COT, albeit without significance between-group differences when compared with NPPV. There was no intergroup difference in mortality between the use of any two of the three respiratory support devices. Several studies in those systematic reviews compared HFNC use with COT and NPPV with COT, although a few studies have compared HFNC use with NPPV. Small sample sizes possibly affected the results of the abovementioned systematic reviews. To overcome these limitations, we performed a systematic review and network meta-analysis (NMA) to compare the effectiveness of three supplemental respiratory support devices in studies that compared at least two of the three techniques (COT, NPPV, and HFNC use) in patients with ARF.

Methods
Protocol and registration
This systematic review was designed according to the Preferred Reporting Items for Systematic review and Meta-Analyses extension statement for reviews incorporating network meta-analyses (details in e-Table 1 in Additional file 1) [23], and the protocol is registered with PROSPERO (CRD42020139105).

Eligibility criteria
Type of studies
We included all randomized-controlled trials (RCTs) reported in English and Japanese regardless of publication status (published, unpublished, and academic abstracts). Randomized crossover, cluster-randomized, and quasi-experimental trials were excluded.

Type of participants
This review included adults (age ≥ 16 years) with acute hypoxic respiratory failure, defined by any of the following criteria: ratio of arterial oxygen partial pressure to fractional inspired oxygen (P/F ratio) < 40.00 kPa; SaO2 or SpO2 < 94% on room air or > 95% with > 6 L/min; and PaO2 < 8.00 kPa with room air or < 10.67 kPa with O2. This meta-analysis excluded studies in which more than half of the patients had post-extubation respiratory failure, acute exacerbation of chronic obstructive pulmonary disease (COPD), acute exacerbation of asthma, hypercapnia (> 6.00 kPa), tracheostomy, post-surgical status, trauma, and do-not-resuscitate orders. The exclusion criteria were limited to factors that were judged clinically appropriate for exclusion by the participating clinicians.

Types of interventions and comparators
We included RCTs comparing two of the following three methods before tracheal intubation:

1. COT: Low-flow nasal cannula, face mask, and venturi mask (with no limit on the flow rate).
2. NPPV: The type of mask and mode, duration of ventilation, and methods of weaning were not limited.
3. HFNC: The flow rate and FIO2 were not limited.

Type of outcomes
The outcome measures included a primary outcome of short-term mortality at the end of the follow-up period (< 90 days), ICU discharge, and hospital discharge. The secondary outcome was the rate of intubation during ICU stay.

Information sources
We searched for eligible trials in the following databases: The Cochrane Central Register of Controlled Trials (CENTRAL); MEDLINE via PubMed; EMBASE; and Ichushi, a database of Japanese research papers. Additionally, we searched for ongoing trials in The World
Health Organization International Clinical Trials Platform Search Portal. For cases with unknown data, the authors were contacted.

**Search**
We used the search terms “ARDS”, “adult respiratory distress syndrome”, “respiratory failure”, or “acute lung injury” AND “non-invasive ventilation”, “NPPV”, “oxygen therapy”, “HFNC”, or “high-flow therapy” in searches performed in December 2020 (details in e-Table 2 in Additional file 1).

**Study selection**
Two of the three physicians (TM, HO, and HY) screened the title and abstract or the full text at the first and second screenings, respectively, for relevant studies and independently extracted data from the included studies into standardized data forms. Disagreements, if any, were resolved by discussion with one of three physicians who did not screen that particular study; original authors were contacted for clarification as required. For abstract-only studies that could not be evaluated for eligibility based on our review criteria, we attempted to contact the authors. Discrepancies between two reviewers were resolved by mutual discussion or discussion with a third reviewer as needed.

**Data collection process and data items**
After identifying studies in the second screening, data were extracted from each study by the reviewers (TM, HO, and HY) using two tools: the Cochrane Data Collection Form (RCTs only) [24] and Review Manager (RevMan) software V.5.3.5 (Cochrane Collaboration) [25]. We extracted the following study characteristics:

1. Methods: study design, total study duration, number and locations of study centers, study setting, withdrawals, and date of study initiation
2. Participants: number, mean age, age range, sex, severity of condition, diagnostic criteria, and inclusion/exclusion criteria
3. Interventions: treatment approaches and comparison methods
4. Outcomes: primary and secondary outcomes that were specified and collected, and the timepoints reported

**Risk of bias within individual studies**
The risk of bias of primary outcomes in the included studies was independently assessed by two of the three authors (TM and HO) using the Cochrane Risk of Bias tool 1.0 (Cochrane Collaboration) [26, 27] in seven domains: (a) random sequence generation, (b) allocation concealment, (c) blinding of participants and personnel, (d) blinding of outcome assessors, (e) incomplete outcome data, (f) selective outcome reporting, and (g) other sources of bias. The risk of each bias was graded as low, unclear, or high. Discrepancies between the two reviewers were resolved through discussion among themselves or with a third reviewer as necessary.

**Statistical analyses**

**Direct comparison meta-analysis**
A pairwise meta-analysis was performed using RevMan 5.3 (RevMan 2014). Forest plots were used for meta-analysis, and the effect size was expressed as relative risk (RR) with the 95% confidence interval (CI) for categorical data and as weighted mean differences with the 95% CI for continuous data. Outcome measures were pooled using a random effect model for study-specific effects in measures. For all analyses, a two-sided p value < 0.05 was considered statistically significant.

Study heterogeneity between trials for each outcome was assessed by visually inspecting forest plots and with an I² statistic to quantify inconsistency [28] (RevMan; I² = 0–40, 40–60, 60–80, and 80–100% indicated minimal, moderate, substantial, and considerable heterogeneity, respectively). When heterogeneity was identified (I² > 50%), we investigated the reason and quantified it using the Chi-square test.

We planned to use a funnel plot, Begg’s adjusted rank correlation test, and Egger’s regression asymmetry test to investigate publication bias if ≥ 10 studies were available (RevMan) [29]. As < 10 studies were included, we did not test for funnel plot asymmetry.

**Network comparison meta-analysis**

**Data synthesis**
A network plot was constructed to determine the number of studies and patients included in this meta-analysis. An NMA using netmeta 0.9-5 R-package (version 3.5.1) was performed via a frequentist-based approach with multivariate random-effects meta-analysis, and effect size was expressed as the RR (95% CI). Covariance between two estimates from the same study shows the variance of data in the shared arm, as calculated in a multivariable meta-analysis performed using the GRADE Working Group Approach for an NMA [30, 31].

**Transitivity**
The transitivity assumption underlying the NMA was evaluated by comparing the distribution of clinical and methodological variables that could act as effect modifiers across treatment comparisons.

**Ranking**
Ranking plots (rankograms) were constructed based on the probability that a given treatment had the highest event rate for each outcome. The surface under
the cumulative ranking curve (SUCRA), which is a simple transformation of the mean rank, was used to determine treatment hierarchy [32] and was constructed using standard software (Stata 15.0, Stata, TX, USA).

**Risk of bias across studies** The assessment of the risk of bias across studies followed considerations of pairwise meta-analysis, and conditions associated with “suspected” and “undetected” across-study bias were determined by the presence of publication bias on a direct comparison.

**Indirectness** We evaluated the indirectness, classified as “no”, “some”, or “major” concern, of each study included in the NMA based on its relevance to the research question, including the study population, interventions, outcomes, and study setting. The study-level judgments could be combined with the percentage contribution matrix.

**Imprecision** The approach to imprecision involved comparing the range of treatment effects included in the 95% CI with the range of equivalence. We assessed the heterogeneity of treatment effects for a clinically important risk ratio of < 0.8 or > 1.25 in the CIs.

**Heterogeneity** To assess the level of heterogeneity, we compared the posterior and predictive distributions of the estimated heterogeneity variance [33]. Concordance between assessments based on CIs and prediction of intervals, both of which do not capture heterogeneity, were used to assess the importance of heterogeneity of treatment effects for a clinically important risk ratio of < 0.8 or > 1.25 in prediction intervals.

**Assessment of inconsistency** The inconsistency of the network model was estimated using inconsistency factors and their uncertainties. We statistically evaluated consistency using the design-by-treatment interaction test [34]. For inconsistency, p values less than 0.05, between 0.05 and 0.10, and otherwise were classified as “Major concerns”, “Some concerns”, and “No concerns”, respectively.

**Additional analyses** If there were sufficient data, we conducted a subgroup analysis of the severity of respiratory failure (P/F < 200) and the cause of respiratory failure (immunocompromised patients, excluding congestive heart failure (CHF)/acute exacerbation of COPD patients) to investigate the impact of risk of bias and assess the heterogeneity in participants in each study.

**Results**

**Study selection**

The comprehensive search yielded 5507 records (e-Fig 1 in Additional file 1), of which 27 studies were included in this NMA [9, 35–60]. These 27 studies included two three-group studies that directly compared NPPV with HFNC use and COT. The final analysis included 19, 7, and 5 studies that compared NPPV with COT, HFNC use with COT, and HFNC use with NPPV, respectively. The network structures of each outcome are shown in Fig. 1a and b.

**Study characteristics**

The protocols and characteristics of each study included in the final dataset of the meta-analysis are summarized in Table 1. The quantitative analysis included 4618 patients. The main cause of acute respiratory failure and the severity of respiratory failure are shown in e-Table 1.
Table 1: Study populations, protocols, and characteristics

| Author, year | Sample size | Study setting | Intervention setting | Control setting | Outcomes | Baseline characteristics | The main causes of acute respiratory failure | Age, years | PaO₂:FiO₂ |
|-------------|-------------|---------------|----------------------|-----------------|----------|-------------------------|-------------------------------------------|-----------|-----------|
| Bersten 1991 [51] | 40 | NPPV | COT | 1. Mortality (in-hospital) 2. Intubation | CPE (CHF [56.4%]) | | NPPV: 76 (6) COT: 75 (6) | NPPV: 138 (32) COT: 136 (44) |
| Wysocki 1995 [36] | 41 | NPPV | COT | 1. Mortality (in-ICU) 2. Intubation | AHRF (CHF [30%]) | | NPPV: 64 (18) COT: 62 (11) | NPPV: 191 (94) COT: 170 (77) |
| Antonelli 2000 [49] | 40 | NPPV | COT | 1. Mortality (in-ICU) 2. Intubation | ARDS (patients who underwent solid organ transplantation [100%]) | | NPPV: 45 (19) COT: 44 (10) | NPPV: 142 (29) COT: 149 (22) |
| Delclaux 2000 [56] | 123 | NPPV | COT | 1. Mortality (in-ICU, in-hospital) 2. Intubation | CPE (infection [49.6%]) | | NPPV: 191 (94) COT: 170 (77) | NPPV: 142 (29) COT: 149 (22) |
| Masip 2000 [39] | 37 | NPPV | COT | 1. Intubation | CPE (myocardial infarction [29.7%]) | | NPPV: 75 (11) COT: 79 (5) | NA |
| Delclaux 2000 [56] | 123 | NPPV | COT | 1. Mortality (in-ICU, in-hospital) 2. Intubation | CPE (infection [49.6%]) | | NPPV: 191 (94) COT: 170 (77) | NPPV: 142 (29) COT: 149 (22) |
| Levitt 2001 [41] | 38 | NPPV | COT | 1. Intubation | CHF | | NPPV: 67 (15) COT: 69 (15) | NA |
| Ferrer 2003 [54] | 105 | NPPV | COT | 1. Mortality (in-ICU, 90 days) 2. Intubation | AHRF (Pneumonia [32.4%]) | | NPPV: 61 (17) COT: 62 (18) | NPPV: 102 (21) COT: 103 (23) |
| L’Her 2004 [44] | 89 | NPPV | COT | 1. Mortality (48 h, in-hospital) | CPE (respiratory tract infection [33.7%]) | | NPPV: 84 (6) COT: 84 (6) | NPPV: 157 (71) COT: 167 (73) |
| Park 2004 [38] | 80 | NPPV | COT | 1. Mortality (in-hospital, 15 days, 60 days) 2. Intubation | CPE (myocardial ischemia [37.5%]) | | NPPV: 64 (15) COT: 65 (15) | NPPV: 141 (24) COT: 136 (23) |
| Gray 2008 [53] | 1069 | NPPV | COT | 1. Mortality (7 days, 30 days) 2. Intubation | CPE (ischemic heart disease [17.6%]) | | NPPV: 77 (10) COT: 79 (9) | NA |
| Cosentini 2010 [55] | 47 | NPPV | COT | 1. Mortality (in-hospital) 2. Intubation | CAP [100%] | | NPPV: 65 (17) COT: 72 (13) | NPPV: 249 (25) COT: 246 (20) |
| Qingyuan 2012 [35] | 40 | NPPV | COT | 1. Mortality (in-ICU, in-hospital) 2. Intubation | ALI (immunocompromised patients [30%]) | | NPPV: 44 (14) COT: 49 (14) | NPPV: 225 (17) COT: 234 (27) |
| Elena 2013 [37] | 80 | NPPV | COT | 1. Intubation | AHRF (pneumonia [100%]) | | NA | NA |
| Brambilla 2014 [52] | 81 | NPPV | COT | 1. Intubation | Pneumonia [100%] | | NPPV: 65 (16) COT: 70 (16) | NPPV: 134 (32) COT: 148 (44) |
| Azevedo 2015 [47] | 30 | NPPV | HFNC | 1. Intubation | AHRF (CHF [43%]) | | NA | NA |
| Frat 2018 [50] | 313 | NPPV/HFNC | COT | 1. Mortality (in-ICU, 90 days) 2. Intubation | AHRF | | NPPV: 61 (15) COT: 69 (15) | NPPV: 149 (17) COT: 150 (17) |
Risk of bias within studies
Additional file 1 e-Fig 2 shows the risk of bias; although all studies did not blind their participants and clinicians to the intervention, the risk of bias in the other domains was low (e-Fig 2 in Additional file 1). Finally, all included studies were judged as having

hypoxic respiratory failure was pneumonia, followed by cardiopulmonary edema from CHF. Seven of the 27 included studies comprised many immunocompromised patients. The reported average P/F at randomization differed among the studies (range 95–249).
Network meta-analysis
The results of pairwise comparisons are shown in e-Fig 3 in Additional file 1 (short-term mortality) and e-Fig 4 in additional file 1 (Intubation). Additional file 1 e-Fig 5 shows the funnel plots of each outcome.

Short-term mortality
In the analysis of short-term mortality (including 20 studies), compared with COT, NPPV (RR, 0.88 [95% CI, 0.76–1.01]; low certainty) and HFNC use (RR, 0.32 [95% CI, 0.80–1.08]; low certainty) showed trends for lower mortality risk (Fig. 2a), and no significant difference was observed between NPPV and HFNC use for mortality (RR, 0.94 [95% CI, 0.78–1.15]; low certainty). Anticipated absolute effects and 95% CI between each of the two comparisons decreased by 28 per 1000 (95% CI, −57 to +2) in NPPV vs. COT, by 21 per 1000 (95% CI, −61 to +24) in HFNC use vs. COT, and by 9 per 1000 (95% CI, −35 to +24) in NPPV vs. HFNC use (Table 2).

Confidence in the RR of each comparison and short-term mortality assessed by the GRADE system is shown in Table 3. Incoherence between direct and indirect RRs was observed for all three comparisons determined by p values of inconsistency. All comparisons (NPPV vs. COT, HFNC use vs. COT, and HFNC use vs. NPPV) showed “Major” concerns. The heterogeneity of all three comparisons resulted in “Major” concern outcomes due to the 95% CI of the predicted risk ratio.

A ranking analysis revealed that the hierarchy for efficacy in reducing short-term mortality was HFNC use (SUCRA 72.1), followed by NPPV (SUCRA 68.8) and ultimately, COT (SUCRA 9.0) (Fig. 3a). The summary of findings of the NMA for short-term mortality is shown in Table 2. The estimate and certainty of the evidence of direct, indirect, and network comparisons are summarized in e-Table 3 in additional file 1.

Endotracheal intubation
Twenty-six studies were included in the analysis of endotracheal intubation. Compared with COT, NPPV (RR, 0.81 [95% CI, 0.72–0.91]; low certainty) and HFNC use (RR, 0.78 [95% CI, 0.68–0.89]; low certainty) were associated with statistically significant lower risks of endotracheal intubation (Fig. 2b), while no significant difference was observed between NPPV and HFNC use in the association with endotracheal intubation (RR, 1.04 [95% CI, 0.88–1.22]; low certainty). Anticipated absolute effects (95% CI) between each of the two comparisons decreased by 57 per 1000 (95% CI, −83 to −27) in
Table 2 Summary of findings of network meta-analysis for short-term mortality

| Estimates of effects, credible intervals, and certainty of the evidence for oxygen therapy in septic patients with acute respiratory failure. | Frequency NMA-Sof table |
|---|---|
| **BENEFITS** | |
| Patients or population: Septic patients with acute respiratory failure who need oxygen therapy | |
| Interventions: One of the following oxygen therapies: NPPV, HFNC, or COT | |
| Comparator (reference): One of the other therapies other than the therapy included in intervention | |
| Outcome: Short-term mortality | |
| Setting: In-hospital | |

| Total studies: 20 | Relative effect (95% CI)** |
|---|---|
| Total Patients: 4,014 | Anticipated absolute effect (95% CI)*** |
| Without intervention | With intervention | Difference | Certain of the evidence | Ranking**** (SU CRA) | Interpretation of Findings |
| NPPV | (15 RCT; 2,559 participants) | 0.88 | (0.76 to 1.01) | Network estimate | 237 per 1000 | 209 per 1000 | 28 fewer per 1000 (57 fewer to 2 more) | ⚠️ | 1.6 | (68.8) | Probably superior |
| HFNC | (6 RCT; 1,509 participants) | 0.93 | (0.86 to 1.08) | Network estimate | 303 per 1000 | 282 per 1000 | 21 fewer per 1000 (61 fewer to 24 more) | ⚠️ | 1.6 | (72.1) | Probably superior |
| ▢ COT | Comparator | Reference | No estimable | No estimable | No estimable | Reference | Comparator | (9.0) | Comparator |

NPPV vs. COT and 70 per 1000 (95% CI, –101 to –35) in HFNC use vs. COT, and increased by 9 per 1000 (95% CI, –28 to +51) in NPPV vs. HFNC use (Table 4).

Confidence in the RR of each comparison and intubation assessed according to the GRADE system (Table 3) showed incoherence between direct and indirect RRs for all three comparisons determined by the p value of inconsistency. All comparisons (NPPV vs. COT, HFNC use vs. COT, and HFNC use vs. NPPV) showed “Major” concerns. The heterogeneity of all three comparisons resulted in “Major” concern due to the 95% CI of the predicted risk ratio. The ranking analysis revealed that the hierarchy for efficacy in reducing intubation was HFNC use (SUCRA 79.4), followed by NPPV (SUCRA 70.2),
and ultimately, COT (SUCRA 0.4) (Fig. 3b). Table 4 in additional file 1 summarizes the NMA findings for intubation; e-Table 3 summarizes the estimate and certainty of the evidence of direct, indirect, and network comparisons.

### Results of additional analyses

Of the 27 RCTs included in this study, only 14 focused on a single cause: 5, immunocompromised status; 4, pneumonia; 3, CHF; 1, post-transplant solid tumors; and 1, interstitial pneumonia. Therefore, the sensitivity analyses performed included the subgroup analysis of the severity of respiratory failure (P/F < 200) and the cause of respiratory failure (immunocompromised patients, excluding CHF/acute exacerbation of COPD patients). In all subgroup analyses, both short-term mortality and intubation rates were similar to those in the main analysis (e-Table 4 in Additional file 1).

### Discussion

This systematic review and NMA showed that NPPV and HFNC use were associated with lower risks of intubation compared with COT rather than an improved mortality risk. The SUCRA values of intubation for HFNC use and NPPV in the NMA showed similar effects to those for NPPV and HFNC use. These results are consistent with those of previous systematic reviews (non-NMA) [18, 21, 61]. A systematic review by Zhao et al. [21] that included 11 studies (n = 3459) compared HFNC use with COT or NPPV and found that unlike NPPV, HFNC use reduced the intubation rate compared with COT. Those studies differ from this study as they included many post-extubation studies. Despite the results of the previous review being consistent with those of our study—that HFNC use reduces the intubation rate compared with COT—the inclusion of studies examining the prevention of re-intubation after extubation would result in a deviation from the clinical question in this study. The sample size was inadequate due to the lack of RCTs that directly compared NPPV and HFNC use, and previous reviews have not shown a significant difference between NPPV and HFNC use. To complement the limitations of the existing studies, a systematic review using an NMA was necessary.

This NMA is the second study to demonstrate the effectiveness of NPPV and HFNC use in ARF. Ferreyro et al. [62] first reported an NMA describing the effects of noninvasive oxygenation strategies (e.g., NPPV and HFNC) on short-term mortality and intubation rates.
HFNC use) for patients with acute hypoxic respiratory failure. They concluded that treatment with noninvasive respiratory support devices was associated with a low mortality risk compared with standard oxygen therapy. Although the results of the intubation rate in this NMA are similar to those in the NMA by Ferreyro et al., the effect of NPPV and HFNC use on mortality, compared with that of COT, differed from the results in Ferreyro et al.’s study. In this NMA, we found no significant differences in the mortality risk between NPPV or HFNC use and COT.

Differences in the number of studies included in the NMA due to differences in the study inclusion criteria may have influenced the differences in the results of the two NMAs. First, this NMA included a large proportion of patients with CHF. Ferreyro et al. excluded studies in which patients with CHF constituted the majority of the study population. The clinical presentations of the cases of pneumonia and CHF are often complicated, with pneumonia being reported as a precipitating factor in CHF. Therefore, excluding studies that had patients with CHF may have affected the results of Ferreyro et al.’s NMA. Second, they included studies wherein patients with COPD constituted < 50% of the population while excluding studies that had a majority (> 50%) of patients with COPD. As noninvasive oxygenation strategies are useful in COPD [63], these studies with populations mostly consisting of patients with COPD may have influenced the NMA results. Additionally, the inclusion of patients with COPD increased the heterogeneity of the

![Image](https://example.com/image.png)

**Fig. 3** Surface under the cumulative ranking of each noninvasive oxygen strategies for short-term mortality and intubation: **a** short-term mortality, **b** intubation.
study population. Third, the NMA by Ferreyro et al. included the studies of ARF that occurred after abdominal surgery and chest trauma-associated respiratory failure, while excluding studies of patients after major cardiovascular surgery. Following abdominal surgery, diaphragmatic dysfunction and decreased vital lung capacity can cause atelectasis, resulting in hypoxemic respiratory failure (HRF). However, in pneumonia, which is the main cause of acute HRF, respiratory failure is caused by the decreased functional residual capacity due to inflammatory leachate in the alveoli and ventilator-perfusion mismatch. Analyzing these distinct pathogeneses of respiratory failure in a similar way is problematic and will affect the interpretation of results.

Implications
For patients, respiratory management without intubation is obviously more comfortable. Despite the superiority of

| Table 4 Summary of findings of network meta-analysis for intubation |
|------------------------|------------------|--------------------------|-----------------|-----------------|-----------------|
| Study Population       | Interventions    | Comparator               | Outcome          | Setting          |
| NMA by Ferreyro et al. | Abdominal surgery and chest trauma | One of the following oxygen therapies: NPPV, HFNC, or COT | Intubation         | In hospital      |

| Total studies: 26 Total Patients: 4,507 | Relative effect (95% CI)** | Anticipated absolute effect (95% CI)*** | Certainty of the evidence | Ranking**** (SUCRA) | Interpretation of Findings |
|----------------------------------------|-----------------------------|----------------------------------------|---------------------------|---------------------|---------------------------|
| NPPV (18 RCT, 2,706 participants)      | 0.81 (0.72 to 0.91) Network estimate | 298 per 1000 241 per 1000 57 fewer per 1000 (83 fewer to 27 fewer) | ⬤⬤⬤⬤ Low* | 1.6 (70.2) | Probably superior |
| HFNC (7 RCT, 1,609 participants)       | 0.78 (0.68 to 0.89) Network estimate | 316 per 1000 247 per 1000 70 fewer per 1000 (101 fewer to 35 fewer) | ⬤⬤⬤⬤ Low* | 1.4 (79.4) | Superior |
| COT                                    | Reference comparator         | No estimable                          | No estimable             | Reference comparator | 3.0 (8.4) | Reference comparator |
| NPPV (5 RCT, 584 participants)         | 1.04 (0.88 to 1.22) Network estimate | 230 per 1000 239 per 1000 9 more per 1000 (28 fewer to 51 more) | ⬤⬤⬤⬤ Low* | - | Probably equivalent |
| HFNC                                   | Reference comparator         | No estimable                          | No estimable             | Reference comparator | - | Reference comparator |

NMA, network meta-analysis; NPPV, non-invasive positive pressure ventilation; HFNC, high-flow nasal cannula; COT, conventional oxygen therapy; SOF, summary of findings; SUCRA, surface under the cumulative ranking.

NMA-SoF table definitions
* Lines represent direct comparisons
** Estimates are reported as risk ratio. CI: confidence interval.
*** Anticipated absolute effect. Anticipated absolute effect compares two risks by calculating the difference between the risks of the intervention group with the risk of the control group.
**** Rank for efficacy outcome are presented. Rank statistics is defined as the probabilities that a treatment out of n treatments in a network meta-analysis is the best, the second, the third and so on until the least effective treatment.

GRADE Working Group grades of evidence (or certainly in the evidence)
High quality: We are very confident that the true effect lies close to that of the estimate of the effect.
Moderate quality: we are moderately confident in the effect estimate: The true effect is likely close to the estimate of the effect, but there is a possibility that it is substantially different
Low quality: Our confidence in the effect estimate is limited: The true effect may be substantially different from the estimate of the effect.
Very low quality: We have very little confidence in the effect estimate: The true effect is likely to be substantially different from the estimate effect

Explanatory Footnotes
a) P value of inconsistency was below 0.05.
HFNC use and NPPV being inconclusive in our study, noninvasive respiratory management is useful because per 1000 patients, utilizing NPPV and HFNC will help avoid intubation in 57–70 patients compared with COT. IMV is associated with various adverse events (e.g., VILI and VAP) and needs specific skills for comfortable management. Respiratory management that avoids tracheal intubation can reduce such complications, and the patient is relieved of the intubation discomfort. For the hospital manager, the costs of respiratory management may decrease due to the lower rate of intubation as daily ventilation costs increase healthcare costs by 59% compared with non-ventilation costs [64].

Limitations
This study had several limitations. First, there may be heterogeneity among the studies included, which may have affected the results. The NMA assumption is that the individual trials enrolled similar populations, and the intervention protocol was similar across different studies. Statistical heterogeneity is affected by a consequence of clinical or methodological diversity. Although statistical heterogeneity was a major concern in the mortality and intubation results, the clinical heterogeneity of diseases, including heart failure, pneumonia, and COPD, which were evaluated in this NMA, may not be considered high because of the difficulty in distinguishing between those diseases in the early stages of real-world clinical practice. The studies included in this NMA could not be clearly categorized by the cause of acute hypoxic respiratory failure because they did not include patients with a single cause. Therefore, sensitivity and subgroup analyses could not be performed.

Second, all RCTs included had high risks of performance bias due to the dramatic differences between HFNC use, COT, and NPPV, which made blinding impossible.

Third, although the statistical analysis was based on the assumption that there was no effect modifier, the inclusion of patients with various degrees of respiratory failure and different outcomes may have influenced the results due to undetectable effect modifiers. However, even for outcomes with different baseline risks, the relative effects of the interventions can remain consistent [65]. A sensitivity analysis based on respiratory failure severity should have been conducted; however, it was not performed because of the possibly similar severity (e.g., P/F ~ 200) of the patients in most studies included. Lastly, a few studies compared NPPV and HFNC use, and thus, the sample size was insufficient to compare NPPV and HFNC use. However, a trend in the direction of a difference is present, which may be further evaluated in the future if more studies compare these two noninvasive oxygenation strategies.

Conclusions
The results of this NMA show that both NPPV and HFNC use are associated with lower risks of endotracheal intubation; however, no significant differences in short-term mortality exist between these respiratory support devices.

Abbreviations
ARF: Acute respiratory failure; CENTRAL: Cochrane central register of controlled trials; CI: Confidence interval; COPD: Chronic obstructive pulmonary disease; COT: Conventional oxygen therapy; HFNC: High-flow nasal cannula; HRF: Hypoxemic respiratory failure; ICU: Intensive care unit; IMV: Invasive mechanical ventilation; NMA: Network meta-analysis; NPPV: Noninvasive positive-pressure ventilation; P/F ratio: Fractional inspired oxygen; RCT: Randomized controlled trial; RevMan: Review manager; RR: Relative risk; SUCRA: Surface under the cumulative ranking curve; VAP: Ventilator-associated pneumonia; VILI: Ventilator-induced lung injury

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Authors’ contributions
HY participated in designing the study, acquiring data, performing statistical analyses, interpreting the data, and drafting the manuscript. HO conceived the study and participated in the interpretation of data, acquisition of data, and drafted the manuscript. MN conceived the study, participated in the designing of the study and interpretation of data, and drafted the manuscript. NS conceived the study, participated in designing the study and interpreting the data, and drafted the manuscript. The authors read and approved the final manuscript.

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The datasets generated during and/or analyzed during the current study are not publicly available due to post hoc analyses by co-authors but are available from the corresponding author on reasonable request.
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Competing interests
The authors declare that they have no competing interests.

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References
1. Ware LB, Matthy MA. The acute respiratory distress syndrome. N Engl J Med. 2000;342:1334–49.
2. Cook D, Brower R, Cooper J, Brochard L, Vincent JL. Multicenter clinical research in adult critical care. Crit Care Med. 2002;30:1636–43.
3. Mattison G, Altobelli N, Kacmarek RM. Noninvasive positive-pressure ventilation in acute respiratory failure outside clinical trials: experience at the Massachusetts General Hospital. Crit Care Med. 2008;36:441–7.
4. Esteban A, Anzueto A, Frutos F, Alia I, Brochard L, Stewart TE, Benito S, Epstein S, Apesteguía C, Nightingale P, Aroniga AG, Tobin MJ. Mechanical Ventilation International Study Group. Characteristics and outcomes in adult patients receiving mechanical ventilation: a 28-day international study. JAMA. 2002;287:3455–5.
5. Kollef MH. What is ventilator-associated pneumonia and why is it important? Respir Care. 2005;50:714–21.
6. Boles JM, Bion J, Connors A, Herridge M, Marsh B, Melot C, Pearl R, Silverman H, Stanchina M, Veillard-Baron A, Welte T. Weaning from mechanical ventilation. Eur Respir J. 2007;29:1033–56.
7. Demoule A, Girou E, Richard JC, Taille S, Brochard L. Increased use of noninvasive ventilation in French intensive care units. Intensive Care Med. 2006;32:1747–55.
8. Thille AW, Richard JC, Brochard L. The decision to extubate in the intensive care unit. Am J Respir Crit Care Med. 2012;187:1294–302.
9. Frat JP, Thille AW, Mercat A, Girault C, Ragot S, Perbet S, Prat G, Boulain T, Moravcek E, Cottetreau A, Devaquet J, Neef S, Razazi K, Mira JP, Argaud L, Chakarian JC, Ricard JD, Wittebole X, Chevalier S, Herbland A, Fartoukh M, Constantin JN, Tonnellerme J, Pironet M, Mathonnet A, Bédunau G, Deléage-Métrieux C, Richard JC, Brochard L, Robert R. FLORALI Study Group, Network REVA. High-flow oxygen through nasal cannula in acute hypoxic respiratory failure. N Engl J Med. 2015;372:2185–96.
10. Hernández G, Vaquero C, González P, Subira C, Frutos-Vivar F, Rialp G, Laborda C, Colinas L, Cuerda R, Fernández R. Effect of postextubation high-flow nasal cannula vs conventional oxygen therapy on reintubation in low-risk patients: a randomized clinical trial. JAMA. 2016;315:1359–61.
11. Zhang Y, Fang C, Dong BR, Wu T, Deng IL. Oxygen therapy for pneumonia in adults. Cochrane Database Syst Rev. 2012;3:CD006607.
12. Xu XP, Zhang XC, Hu SL, Xu JY, Xie JF, Liu SQ, Liu L, Huang YZ, Guo FM, Yang Y, Qu HB. Noninvasive ventilation in acute hypoxic nonhypercapnic respiratory failure: a systematic review and meta-analysis. Crit Care Med. 2017;45:e727–33.
13. Bello G, De Pascale G, Antonelli M. Noninvasive ventilation. Clin Chest Med. 2016;37:711–21.
14. Hill NS. Complications of noninvasive ventilation. Respir Care. 2000;45:480–1.
15. Rochweerg B, Granton D, Wang DX, Helvitz Y, Einav S, Frat JP, Mekontso-Desap A, Schreiber A, Azoulay E, Mercat A, Demoule A, Lemiaye V, Pesenti A, Riviello ED, Mauri T, Mancebo J, Brochard L, Burns K. High-flow nasal cannula compared with conventional oxygen therapy for acute hypoxic respiratory failure: a systematic review and meta-analysis. Intensive Care Med. 2019;45:563–72.
16. Xu Z, Li Y, Zhou J, Li X, Huang Y, Liu X, Burns KEA, Zhong N, Zhang H. High-flow nasal cannula oxygen therapy on mortality and intubation rate in acute respiratory failure: a systematic review and meta-analysis. Crit Care Med. 2017;45:e449–56.
17. Ny VN, Luo J, Yu H, Liu D, Liang BM, Yao R, Liang ZA. Can high-flow nasal cannula reduce the rate of reintubation in adult patients after extubation? a meta-analysis. BMC Pulm Med. 2017;17:142.
18. Afraimi MA, Alhmari MD, Abtabbi H, Alfalaseh S, Albasvari I, Mubasher M. Evaluation of efficacy of non-invasive ventilation in non-COPD and non-trauma patients with acute hypoxic respiratory failure: a systematic review and meta-analysis. Ann Thorac Med. 2015;10:16–24.
19. Zhu Y, Yin H, Zhang R, Wei J. High-flow nasal cannula oxygen therapy versus conventional oxygen therapy in patients with acute respiratory failure: a systematic review and meta-analysis of randomized controlled trials. BMC Pulm Med. 2017;17:201.
20. Zhao H, Wang H, Sun F, Lye S, An Y. High-flow nasal cannula therapy is superior to conventional oxygen therapy but not to noninvasive mechanical ventilation on intubation rate: a systematic review and meta-analysis. Crit Care. 2017;21:184.
21. Ny VN, Luo J, Yu H, Liu D, Liang BM, Liang ZA. The effect of high-flow nasal cannula in reducing the mortality and the rate of endotracheal intubation when used before mechanical ventilation compared with conventional oxygen therapy and noninvasive positive pressure ventilation. A systematic review and meta-analysis. Am J Emerg Med. 2018;36:226–33.
22. Hutton B, Salanti G, Caldwell DM, Chaimani A, Schmid CH, Cameron P, Ioannidis JPA, Straus S, Thahir K, Jansen JP, Mouliv N, Catalá-López F, Gatzics PC, Dickersin K, Bouton I, Altman DG, Moher D. The PRISMA extension statement for reporting of systematic reviews incorporating network meta-analyses of health care interventions: checklist and explanations. Ann Intern Med. 2015;162:777–84.
23. TC C. Data collection form-intervention review for RCTs only. Secondary data collection form-intervention review for RCTs only. 2014. https://diplo. cochrane.org/data-extraction-forms (accessed September 8, 2019).
24. TC C. RevMan S Download and installation. Secondary RevMan S download and installation. 2014. https://community.cochrane.org/help/tools-and-sofware/revman-5/revman-5-download/installation (accessed September 8, 2019).
25. TC C. Cochrane handbook for systematic reviews of interventions. 2011. https://training.cochrane.org/handbook (accessed September 8, 2019).
26. TC C. Table 8.5.d: Criteria for judging risk of bias in the ‘Risk of bias’ assessment tool. Secondary Table 8.5.d: Criteria for judging risk of bias in the ‘Risk of bias’ assessment tool. 2011. http://handbook-5-1.cochrane.org (accessed September 8, 2019).
27. Higgins JP, Thompson SG. Quantifying heterogeneity in a meta-analysis. Stat Med. 2002;21:1539–58.
28. Sterne JA, Sutton AJ, Ioannidis JP, Raffin S, Terrin N, Rodriguez M, Gibbons D, Duval S, Altman DG, Begg C, Sterne JA, Kessels AG, Guyatt GH, Group GW. A GRADE Working Group approach for grading the quality of treatment effect estimates from network meta-analysis. BMJ. 2014;349:g6307.
29. Puhan MA, Schunemann HJ, Murad MH, Li T, Brignardello-Petersen R, Singh JA, Kessels AG, Guyatt GH, Group GW. A GRADE Working Group approach for rating the quality of treatment effect estimates from network meta-analysis. BMJ. 2014;349:g6307.
30. Yepez-Nunez JJ, Li SA, Guyatt G, Jack SM, Brozek JL, Beyene J, Murad MH, Rochwerger B, Mbuagbaw L, Zhang Y, Florez JD, Semkin RA, Sadeghirad B, Mustafa R, Santesso N, Schunemann HJ. Development of the summary of findings table for network meta-analysis. J Clin Epidemiol. 2019;115:1–13.
31. Salanti G, Ades AE, Ioannidis JP. Graphical methods and numerical summaries for presenting results from multiple-treatment meta-analysis: an overview and tutorial. J Clin Epidemiol. 2011;64:163–71.
33. Rhodes KM, Turner RM, Higgins JP. Predictive distributions were developed for the extent of heterogeneity in meta-analyses of continuous outcome data. J Clin Epidemiol. 2015;68:52–60.

34. Higgins JP, Jackson D, Barrett JK, Lu G, Ades AE, White IR. Consistency and inconsistency in network meta-analysis: concepts and models for multi-arm studies. Res Synth Methods. 2012;3:98–110.

35. Zhan Q, Sun B, Liang L, Yan X, Zhang L, Yang J, Wang L, Ma Z, Shi L, Wei L, Li G, Yang L, Shi Z, Chen Y, Xu Q, Li W, Zhu X, Wang Z, Sun Y, Zhuo J, Liu Y, Li X, Wang C. Early use of noninvasive positive pressure ventilation for acute lung injury: a multicenter randomized controlled trial. Crit Care Med. 2012; 40:4955–60.

36. Wysocki M, Tric L, Wolff MA, Millet H, Hemm B. Non-invasive positive pressure support ventilation in patients with acute respiratory failure. A randomized comparison with conventional therapy. Chest. 1995;107:761–8.

37. Prina E, Brambilla AM, Alberti S, Piffer F, Bozzani R, Vismont S, Bellini A, Cioni L, Milic-Emili J, Mallo S, Pellegrino R. Non-invasive continuous positive airway pressure versus oxygen Venturi in severe acute respiratory failure due to pneumonia: a randomized controlled trial. European Respiratory Society Annual Congress. 2013;42:P4764.

38. Park M, Sangean MC, Volpe Mde S, Feltrim MIZ, Nozawa E, Leite PF, Amato MBP, Lorenzi-Filho G. Randomized, prospective trial of oxygen, continuous positive airway pressure, and bilevel positive airway pressure by face mask in acute respiratory failure. Crit Care Med. 2004;32:2407–15.

39. Masip J, Bérbésé A, Páez J, Vecilla F, Carriazo R, Padró J, Paz MA, de Oto J, Ballús J. Non-invasive pressure support ventilation versus conventional oxygen therapy in acute respiratory failure due to neurological conditions. A randomised trial. Lancet. 2000;356:2026–32.

40. Makdde O, Monsomboon A, Surabenjawong U, Praphruetkit N, Chaisirin W, Park M, Sangean MC, Volpe Mde S, Feltrim MIZ, Nozawa E, Leite PF, Amato MBP, Lorenzi-Filho G. Randomized, prospective trial of oxygen, continuous positive airway pressure, and bilevel positive airway pressure by face mask in acute respiratory failure. Crit Care Med. 2004;32:2407–15.

41. Levitt MA. A prospective, randomized trial of BiPAP in severe acute respiratory failure: a multicenter randomized trial. Crit Care Med. 2005;33:455–60.

42. Pregnolato P, Bragagnolo C, del Re M, Cantini M, Ferrari G, Tripodi A, Thiel C, De Constanti G, Filippi A, Rella A, Mancini M, Morelli L, Incalzi RA, Ghi S, Malagoni AM, D’Agostino A. Non-invasive ventilation in acute respiratory failure due to exacerbation of chronic obstructive pulmonary disease. Intensive Care Med. 2015;41:192–200.

43. Hématologie (GRRR-OH). Effect of noninvasive ventilation vs oxygen therapy in patients undergoing solid organ transplantation: a randomized trial. Intensive Care Med. 2015;41:192–200.

44. Bersten AD, Holt AW, Vedig AE, Skowronski GA, Baggoley CJ. Treatment of severe cardiogenic pulmonary edema with continuous positive airway pressure delivered by face mask. N Engl J Med. 1991;325:1825–30.

45. Brambilla AM, Alberti S, Prina E, Nicolli F, Forino MD, Nava S, Ferrari G, Corradi F, Pelosi P, Bignamini A, Tarsia P, Cosentino R. Helmet CPAP vs. oxygen therapy in severe hypoxic respiratory failure due to pneumonia. Intensive Care Med. 2014;40:402–9.

46. Gray A, Goodacre S, Newby DE, Masson M, Sampson F, Nicholl J, Trialists CPO. Noninvasive ventilation in acute cardiogenic pulmonary edema. N Engl J Med. 2008;359:142–51.

47. Ferrer M, Esquinas A, Leon M, Gonzalez G, Alarcon A, Torres A. Noninvasive ventilation in severe hypoxic respiratory failure: a randomized clinical trial. Am J Respir Crit Care Med. 2003;168:1388–44.

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