Associations of Dietary Inflammatory Index With Prediabetes and Insulin Resistance

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Background and Aims: Previous studies suggested that dietary inflammatory index (DII) was associated with a variety of adverse health conditions. However, less is known about the role of DII in prediabetes and insulin resistance (IR). Therefore, this study aimed to investigate whether DII is associated with prediabetes and IR in American adults.

Method and Results: DII scores were calculated using the average of two 24-hour dietary recalls. Linear regression models were performed to evaluate the associations of DII with markers of Type 2 diabetes (T2D) risk, and the associations of DII with prediabetes and IR were estimated using logistic regression model. The diet of the participants showed an anti-inflammatory potential, with a mean DII score of −0.14 (range: −5.83 to +5.32). After controlling for multiple potential confounders, DII scores were positively associated with fasting plasma glucose (FPG) (β: 0.009; 95%CI: 0.005 to 0.012), fasting serum insulin (FSI) (β: 0.083; 95%CI: 0.067 to 0.099) and homeostatic model assessment of insulin resistance (HOMA-IR) (β: 0.092; 95%CI: 0.075 to 0.109). Participants in the highest tertile of DII score have increased odds of prediabetes (OR: 1.40; 95%CI: 1.17 to 1.69; P for trend <0.001) and IR (OR: 1.79; 95%CI: 1.49 to 2.14; P for trend <0.001) compared with those in the first tertile of DII score.

Conclusions: This study indicates that DII was positively associated with FPG, FSI, and HOMA-IR, and a more pro-inflammatory diet was related to increased odds of insulin resistant and prediabetes.

Keywords: dietary inflammatory index, insulin resistance, prediabetes, diabetes, NHANES

INTRODUCTION

Currently, there are about 34.2 million people with diabetes in the US, and the prevalence of diabetes is expected to continue to rise in the future (1). Type 2 diabetes (T2D), the most common type of diabetes, is a leading cause of death and poor health and exerts a large and rapidly increasing burden on the US economy (2).
Prediabetes and insulin resistance (IR), major contributing factors in the development of T2D, were characterized by abnormal glucose metabolism (1, 3, 4). To prevent or delay the development of T2D in early stage, a variety of studies were conducted to identify the associations of modifiable factors (e.g., diet, obesity, and smoking) with prediabetes and IR (5–7). As the most concerned and modifiable factor, dietary intake plays a significant role in human health (5, 8, 9). Findings from prior studies have demonstrated that several pro-inflammatory dietary components (e.g., fat and carbohydrate) were associated with increased risk of IR and also other poor health conditions (8–10), while other anti-inflammatory nutrients (e.g., fiber) may have protective effects on health conditions (11). However, it may be difficult to capture the overall effects of diet on health by studying the relationship of single nutrient with diseases. In this context, the dietary inflammatory index (DII) has been proposed to evaluate the inflammatory potential of the overall diet and link diet to inflammation which was a key driver in the progression of prediabetes and IR (12, 13). The DII score is obtained from the diet of an individual based on the pro- and anti-inflammatory properties of the overall dietary compositions, namely, macronutrients, micronutrients, and some other dietary constituents. Previous studies (namely, the studies performed in NHANES) have shown that DII scores were related to a host of health conditions in the general population, namely, hypertension, cancer, and even death (14, 15). However, the associations of DII with prediabetes and IR have received little attention, and several limited studies on the association of DII with markers of Type 2 diabetes (T2D) risk have yielded conflicting results (16–19).

Thus, the main objective of present study was to investigate the associations of the DII with IR and prediabetes in a national representative population in America. The second aim of this study was to investigate the associations between DII and the markers of T2D risk, namely, fasting plasma glucose [FPG], glycohemoglobin [HbA1c], fasting serum insulin [FSI], and homeostatic model assessment of insulin resistance index [HOMA-IR].

**METHODS**

**Study Participants**

The data of current study were extracted from 2007 to 2016 cycles of the National Health and Nutrition Examination Surveys (NHANES), which are ongoing, stratified, multistage probability surveys of the non-institutionalized civilian population residing in the 50 states and districts of Columbia in the United States. The survey consists of questionnaires administered in the home, followed by a standardized health examination in specially equipped mobile examination centers. Beginning in 1999, most data in this nationally representative survey have been released online in 2-year cycles. The NHANES procedures were approved by the National Center for Health Statistics research ethics review board, and the informed consent was obtained from all participants prior to their inclusion in the surveys. The official website provides more detailed information on the methods and protocols of the NHANES (http://www.cdc.gov/nchs/nhanes.htm).

During the 2007–2016 cycles of the NHANES, a total of 30,724 adults constituted the study sample. Due to physiological and pathological factors, namely, pregnancy, diabetes diagnosis, medication use and fasting status, the assessment of the markers of T2D risk or dietary intake might be potentially biased because of the changed dietary habits or other reasons. Therefore, the exclusion criteria of this study were as follows: a) pregnant women (n = 307); b) participants with known diabetes (n = 3,999, self-reported diabetes or use of insulin or oral glucose-lowering medications); c) those with missing data on FPG, HbA1c, and FSI (n = 15,220); d) those with attending the morning examination after fasting <8 h (n = 748); e) those whose dietary recall status is unreliable or does not meet the minimum criteria (n = 1,723); then, 8,727 remained in the study. After further excluding those with missing data on key covariates including sex, educational attainment, race/ethnicity, body mass index (BMI), family poverty income ratio (PIR), hypertension, serum cotinine, and physical activity (n = 801), finally, a total of 7,926 participants (weighted n = 159,386,984) were enrolled in the present study. The results of comparisons among the included participants and those adults without missing data on T2D risk markers and all participants are shown in Table S1. Notably, during the examinations of FPG & FSI in the NHANES, there were a series of exclusion criteria (e.g., fasting status, taking insulin or oral medications for diabetes, refusing phlebotomy), resulting in nearly half of the participants missed FPG and FSI.

**Assessment of DII**

The development and validation of the DII has previously been reported elsewhere (12, 20). Briefly, dietary intake data for each study participant was first linked to a database that provides a reference global daily mean and standard deviation intake for a total of 45 food parameters from 11 populations around the world. A z-score was derived by subtracting the mean of the database and dividing this value by the standard deviation of the parameter. These z-scores were converted to percentile scores and then centering by doubling and subtracting 1 (from −1 to +1 and centered on 0). Afterwards, each centered proportion was multiplied by the corresponding literature-derived inflammatory effect score for each food parameter. Finally, the overall DII score for each individual is the sum of each food parameter-specific DII score (12).

In this study, dietary data used to calculate the DII score was evaluated by the average of two 24-h dietary recalls. In brief, through face-to-face interview (the first 24-h dietary recall interview) and telephone interview (the second 24-h dietary recall interview), individual foods/beverage consumed during a 24-hour time period before the interview was obtained. The United States Department of Agriculture Food and Nutrient Database for Dietary Studies was used to quantify specific nutrients in the reported dietary components (21). In the NHANES, 31 of the 45 food parameters were available for DII calculation: carbohydrates; protein; fiber; fat; cholesterol; saturated, monounsaturated, and polyunsaturated fatty acids; omega-3 and omega-6 polyunsaturated fatty acids; magnesium;
iron; zinc; selenium; niacin; vitamins A, B1, B2, B6, B12, C, D, E; beta carotene; folic acid; alcohol; caffeine; garlic; onion; ginger; and pepper. The unavailable food parameters included eugenol, saffron, trans fat, turmeric, green/black tea, flavan-3-ol, flavonoids, flavonones, anthocyanidins, isoflavones, thyme/ oregano and rosemary. To control for the effect of total energy intake, the energy adjusted DII scores were calculated per 1,000 calories of food consumed (using the energy-standardized version of the world database). Given that a significant association between DII and c-reactive protein (CRP) was reported in prior study of the NHANES (22), higher DII (i.e., more positive) scores tend to indicate more pro-inflammatory diets and more negative values tend to indicate more anti-inflammatory (12).

**Measurements of the Markers of T2D Risk**

FPG was measured by an enzyme hexokinase method. FSI was measured using an immunoassay method. Hba1c from whole blood was measured using the boronate affinity high performance liquid chromatography system. The concentrations of FPG, FSI, and Hba1c were adjusted for differences in laboratory methodology in the NHANES 2007–2016. The IR status was measured by HOMA-IR, and the 75th percentile value of HOMA-IR was used as a cut-off level to define the IR in present study (23). The corresponding cutoff value was 3.475 for the present study. The HOMA-IR scores were computed as follows: HOMA-IR = [fasting insulin (μU/ml) × fasting glucose (mmol/L)]/22.5 (24). Prediabetes was defined by the Hba1c level of 5.7 to 6.4% and/or impaired fasting glucose level (100–125 mg/dl) and/or impaired glucose tolerance (140–199 mg/dl). In the present study, we also defined undiagnosed diabetes by the Hba1c (≥6.5%) and/or fasting glucose (≥126 mg/ dl) and/or impaired glucose tolerance (≥200 mg/dl) (25).

**Study Covariates**

Potential confounders, namely, age, sex, ethnicity, education levels, PIR, total energy intake, serum cotinine concentration, BMI, physical activity, hypertension, and diseases history were ascertained using questionnaires or physiological and biochemical tests. The sex was categorized as male and female. The ethnicity was categorized as non-Hispanic white, non-Hispanic black, Mexican American, other Hispanic, and other race. Educational attainment was classified into lower than high school, high school, and higher than high school. BMI was defined as weight divided by height squared (kg/m²) (26). PIR was calculated by dividing family (or individual) income by the poverty guidelines specific to the survey year. Physical activity, reported using the Global Physical Activity Questionnaire, was computed as the total metabolic equivalent scores (METs, in minutes/week) summed from the recommended MET scores for each activity. Hypertension was defined as a self-reported history of hypertension, self-reported use of blood pressure medications, systolic blood pressure ≥140 mmHg or diastolic blood pressure ≥90 mmHg. Participants were asked by trained interviewers using the computer-assisted personal interviewing system to obtain health conditions including cardiovascular disease (CVD) diagnosis and cancer diagnosis.

**Statistical Analyses**

Categorical variables were expressed as counts (percentages) and were compared using the chi-squared test. Continuous variables were expressed as mean (standard error, SE) if data presented a normal distribution and were compared using analysis of variance (ANOVA), or expressed as medians (25th–75th) if data presented a skewed distribution and were compared by the Wilcoxon rank sum test. FPG, FSI, Hba1c, and HOMA-IR were log-transformed to better approximate a normal distribution. The associations of DII with FPG, FSI, Hba1c, and HOMA-IR were estimated using linear regression model, and the associations of DII with IR and prediabetes were estimated using logistic regression model. Three models were applied in the present study, with adjustment for potential confounders ascertained based on prior publications (27–31). Model 1 was adjusted for energy intake and age. Model 2 was adjusted for energy intake, age, sex, ethnicity, education levels and PIR. Model 3 was further adjusted for serum cotinine concentration, BMI, hypertension, CVD, cancer and physical activity. Notably, participants with undiagnosed diabetes were excluded in models when investigating the relationship between DII and prediabetes. Moreover, to demonstrate the complex interrelationships among the variables of interest, a path analysis with prediabetes status and IR status as the outcomes of interest was carried out and shown in Figure S1.

Since previous studies have reported significant associations of obesity with prediabetes and IR (32, 33), a subgroup analysis was performed to assess the potential modification effect by obesity. Obesity was defined as BMI ≥30 kg/m² for the present study. Thus, the present study subjects were classified into two subgroups (obesity: BMI ≥30 kg/m²; non-obesity: BMI <30 kg/m²). As an interaction term, DII × Obesity, was inserted in the regression models to obtain P-value for interaction. In addition, the restricted cubic spline models were applied to evaluate the dose–response relationships of DII with prediabetes and IR in all the participants and subgroups, with adjustment for abovementioned covariates.

To evaluate the robustness of the associations of DII with IR and prediabetes, we further performed a sensitive analysis with adjustment for white blood cell count to control the circulating level of inflammatory marker. Additionally, we performed another sensitive analysis with strict criteria to define prediabetes (Hba1c (6.0–6.5%) and/or FPG (110–125 mg/dl)). Lastly, because there were participants with undiagnosed diabetes in the present study, we also performed an analysis to investigate the association between DII and undiagnosed diabetes.

Survey weighing procedures accounted for the effects of stratification and clustering used in the NHANES study design. We constructed fasting subsample weights for analysis as appropriate to obtain US nationally representative estimates. All data cleaning and statistical analyses were performed using R software (version 3.6.3) (34).

**RESULTS**

Table 1 presents the characteristics of the study participants according to the tertiles of DII. A total of 7,926 adults (weighted...
n = 159,386,984) were included in the present study. The DII scores of all the participants ranged from −5.83 to 5.32, with a mean of −0.14. Subjects with higher DII scores were younger, males, and more likely to have higher BMI.

Table 2 shows the associations between DII and the markers of T2D risk. The associations of continuous DII scores with FPG, FSI, and HOMA-IR were significant after adjustment for potential confounders, but the relationship between DII and HbA1c was significant only after adjustment for covariates in model 2. After full adjustment for covariates in model 3, participants in the second or the third tertile of DII score had higher levels of FPG, FSI, and HOMA-IR than those in the first tertile of DII score, and the corresponding β coefficients (95% CI) for the highest tertile of DII score were 0.021 (95%CI: 0.013 to 0.028), 0.191 (95%CI: 0.153 to 0.229), and 0.212 (95%CI: 0.172 to 0.252) for FPG, FSI, and HOMA-IR, respectively.

The relationship between DII and prediabetes according to the tertiles of DII score is listed in Table 3. Compared with the individuals in the lowest tertile of DII, the multivariable-adjusted odds ratios (95% CIs) of prediabetes for the second and the third tertiles of DII was 1.26 (95%CI: 1.08 to 1.47) and 1.40 (95%CI: 1.17 to 1.69), respectively (P for trend <0.001). In the results of subgroup analyses, the association between DII and prediabetes was more evident among participants with obesity (OR:1.50; 95% CI:1.04 to 2.15), but there was no significant modification effect of BMI on the DII-prediabetes association (P value for interaction = 0.904).

The association between DII and IR according to the tertiles of DII score is listed in Table 4. An increment of 1 SD in the DII score (i.e., 1.91 units) was associated with a 26% higher odds of IR (OR: 1.26; 95%CI: 1.18 to 1.36) after adjustment for potential covariates. Additionally, we also assessed the relationship between tertiles of DII score and IR. Compared with the individuals in the lowest tertile of DII, the odds of IR increased by 79% among the individuals in the highest tertile of DII score (OR: 1.79; 95%CI: 1.49 to 2.14). In analysis stratified by BMI, participants with obesity in the highest tertile of DII had an 79% higher odds of IR than those with DII score in the lowest tertile (OR:1.79; 95%CI:1.38 to 2.34). Similarly, participants without obesity in the highest DII tertile had a 78% higher odds of IR than those with DII score in the lowest tertile (OR: 1.78; 95%CI: 1.34 to 2.37). However, the modification effect of BMI on the DII-IR association was insignificant (P-value for interaction = 0.653).

Additionally, the dose–response relationships of DII with prediabetes and IR were further assessed by using the restricted cubic splines. Overall, there was a linear dose–response relationship between DII score and IR (P for

### Table 1: Demographic characteristics of adults aged 18 years and older according to the tertiles of dietary inflammatory index (National Health and Nutrition Examination Survey 2007–2016).

| Characteristics | DII |  |  |  |
|-----------------|-----|---|---|---|
|                 | T1 (−5.83, −1.04) | T2 (−1.05, 0.76) | T3 (0.77, 5.32) | P | P for trend |
| Age, years, mean (SE)* | 49.8 (0.5) | 48.2 (0.4) | 40.6 (0.4) | <0.001 | <0.001 |
| BMI, kg/m², mean (SE) | 27.6 (0.2) | 28.8 (0.2) | 29.1 (0.2) | <0.001 | <0.001 |
| BMI <30 | 1858 (70.9) | 1680 (63.6) | 1674 (63.4) | <0.001 | <0.001 |
| BMI ≥30 | 784 (29.7) | 962 (38.4) | 968 (38.6) | <0.001 | <0.001 |
| PIR, mean (SE) | 3.4 (0.1) | 3.0 (0.1) | 2.6 (0.1) | <0.001 | <0.001 |
| Serum cotinine, ng/ml, median (25th–75th) | 0 (0–0.1) | 0 (0–1.5) | 0.2 (0–166) | <0.001 | <0.001 |
| Sex, n(%) | 1,599 (60.5) | 1,352 (51.2) | 1,203 (45.5) | <0.001 | <0.001 |
| Male | 1,043 (39.5) | 1,290 (48.8) | 1,439 (54.5) | <0.001 | <0.001 |
| Female | 1,599 (60.5) | 1,352 (51.2) | 1,203 (45.5) | <0.001 | <0.001 |
| Education levels, n(%) |  | | | <0.001 | <0.001 |
| <high school | 440 (16.7) | 538 (20.4) | 604 (22.9) | <0.001 | <0.001 |
| high school | 433 (16.4) | 648 (24.5) | 888 (33.6) | <0.001 | <0.001 |
| >high school | 1,768 (67) | 1,466 (55.1) | 1,150 (43.5) | <0.001 | <0.001 |
| Ethnicity/Race, n(%) |  | | | <0.001 | <0.001 |
| Non-Hispanic White | 1,253 (47.4) | 1,229 (48.5) | 1,251 (47.4) | <0.001 | <0.001 |
| Non-Hispanic Black | 340 (12.9) | 419 (15.9) | 658 (24.9) | <0.001 | <0.001 |
| Hispanic | 289 (10.9) | 267 (10.1) | 226 (8.6) | <0.001 | <0.001 |
| Mexican American | 364 (13.8) | 460 (17.4) | 357 (13.5) | <0.001 | <0.001 |
| Other Race | 396 (15) | 267 (10.1) | 150 (5.7) | <0.001 | <0.001 |
| Physical activity (MET)*, median (25th–75th) | 1,440 (360–3600) | 1,440 (240–4091) | 1,560 (240–5760) | 0.02 | 0.80 |
| Hypertension, yes, n (%) | 1,020 (38.6) | 971 (36.8) | 877 (33.2) | <0.001 | <0.001 |
| CVD, ever, n (%) | 220 (8.3) | 231 (8.7) | 194 (7.3) | 0.16 | 0.191 |
| Cancer, ever, n (%) | 290 (11) | 213 (8.1) | 162 (6.1) | <0.001 | <0.001 |

*SE, standard error.
**BMI, body mass index.
**PIR, family poverty income ratio.
**Serum cotinine, a metabolite of nicotine, is a marker of exposure to tobacco smoke.
**MET, metabolic equivalents.
**Hypertension was defined as a self-reported history of hypertension, self-reported use of blood pressure medications, systolic blood pressure ≥140 mmHg or diastolic blood pressure ≥90 mmHg.
**CVD, cardiovascular diseases including heart failure, coronary heart disease, angina, heart attack, stroke.
non-linearity = 0.112) and prediabetes (P for non-linearity = 0.342), but a significant non-linear association of DII with IR was observed in the participants without obesity (P for non-linearity = 0.009; Figure 1).

The results of sensitive analyses are shown in Tables S2–S4. Specifically, the significant associations of DII with IR and prediabetes were persisted after adjusting for white blood cell count to control the circulating level of inflammatory marker.

### TABLE 2 | β-Coefficients (95% CIs) for the relationship between the dietary inflammatory index and the markers of T2D risk (National Health and Nutrition Examination Survey 2007–2016).

| Outcomes | β (95% CI) | P for trend |
|----------|------------|-------------|
| FPGc | Per SD increase | T1 (−5.83, −1.04) | T2 (−1.05, 0.76) | T3 (0.77, 5.32) |
| Model 1g | 0.004 (0.001, 0.007) | ref | 0.019 (0.010, 0.027) | 0.012 (0.004, 0.019) | 0.003 |
| Model 2h | 0.010 (0.007, 0.013) | ref | 0.021 (0.014, 0.029) | 0.025 (0.017, 0.033) | <0.001 |
| Model 3i | 0.009 (0.005, 0.012) | ref | 0.017 (0.010, 0.025) | 0.021 (0.013, 0.028) | <0.001 |

FSId | Per SD increase | T1 (−5.83, −1.04) | T2 (−1.05, 0.76) | T3 (0.77, 5.32) |
| Model 1g | −0.000 (−0.003, 0.002) | ref | 0.006 (0.000, 0.012) | 0.000 (−0.005, 0.006) | 0.941 |
| Model 2h | 0.004 (0.002, 0.006) | ref | 0.009 (0.004, 0.014) | 0.010 (0.005, 0.015) | <0.001 |
| Model 3i | 0.002 (−0.003, 0.004) | ref | 0.006 (0.001, 0.011) | 0.005 (0.000, 0.010) | 0.033 |

HbA1C* | Per SD increase | T1 (−5.83, −1.04) | T2 (−1.05, 0.76) | T3 (0.77, 5.32) |
| Model 1g | −0.000 (−0.003, 0.002) | ref | 0.006 (0.000, 0.012) | 0.000 (−0.005, 0.006) | 0.941 |
| Model 2h | 0.004 (0.002, 0.006) | ref | 0.009 (0.004, 0.014) | 0.010 (0.005, 0.015) | <0.001 |
| Model 3i | 0.002 (−0.003, 0.004) | ref | 0.006 (0.001, 0.011) | 0.005 (0.000, 0.010) | 0.033 |

HOMA-IRf | Per SD increase | T1 (−5.83, −1.04) | T2 (−1.05, 0.76) | T3 (0.77, 5.32) |
| Model 1g | 0.104 (0.085, 0.122) | ref | 0.197 (0.150, 0.243) | 0.251 (0.209, 0.294) | <0.001 |
| Model 2h | 0.108 (0.089, 0.126) | ref | 0.188 (0.144, 0.232) | 0.255 (0.207, 0.302) | <0.001 |
| Model 3i | 0.092 (0.075, 0.109) | ref | 0.139 (0.105, 0.174) | 0.212 (0.172, 0.252) | <0.001 |

**aCI, confidence interval.**
**bSD, standard deviation.**
**cFPG, fasting plasma glucose.**
**dFSI, fasting serum insulin.**
**eHbA1c, glycohemoglobin.**
**fHOMA-IR, homeostatic model assessment of insulin resistance.**
**gModel 1 adjusted for energy (continuous) and age (continuous).**
**hModel 2 adjusted for energy (continuous), age (continuous), sex (categorical), education levels (categorical), ethnicity/race (categorical), PIR (continuous).**
**iModel 3 adjusted for energy (continuous), age (continuous), sex (categorical), education levels (categorical), ethnicity/race (categorical), PIR (continuous), BMI (categorical), hypertension (categorical), CVD (categorical), cancer (categorical), serum cotinine (continuous), and physical activity (continuous).**

### TABLE 3 | Multivariable-adjusted odds ratios (95% CIs) of prediabetes according to the tertiles of dietary inflammatory index (National Health and Nutrition Examination Survey 2007–2016).

| Models | OR (95% CI) | P for trend | P for interaction*h |
|--------|-------------|-------------|---------------------|
| Overall | Per SD increase | T1 (−5.83, −1.04) | T2 (−1.05, 0.76) | T3 (0.77, 5.32) |
| Model 1g | 1.25 (1.17, 1.34) | ref | 1.44 (1.25, 1.66) | 1.68 (1.44, 1.97) | <0.001 |
| Model 2h | 1.20 (1.12, 1.29) | ref | 1.32 (1.14, 1.53) | 1.49 (1.25, 1.78) | <0.001 |
| Model 3i | 1.17 (1.09, 1.26) | ref | 1.26 (1.08, 1.47) | 1.40 (1.17, 1.66) | <0.001 |
| BMI ≥30 kg/m² | | | | |
| Model 1g | 1.19 (1.04, 1.35) | ref | 1.51 (1.15, 1.97) | 1.44 (1.05, 1.97) | 0.036 |
| Model 2h | 1.20 (1.04, 1.39) | ref | 1.48 (1.13, 1.95) | 1.46 (1.03, 2.06) | 0.053 |
| Model 3i | 1.21 (1.05, 1.41) | ref | 1.52 (1.14, 2.02) | 1.50 (1.04, 2.15) | 0.044 |

| BMI <30 kg/m² | | | | |
| Model 1g | 1.24 (1.15, 1.34) | ref | 1.33 (1.09, 1.61) | 1.67 (1.38, 2.00) | <0.001 |
| Model 2h | 1.16 (1.07, 1.26) | ref | 1.19 (0.98, 1.45) | 1.41 (1.17, 1.70) | <0.001 |
| Model 3i | 1.16 (1.07, 1.26) | ref | 1.18 (0.97, 1.43) | 1.39 (1.15, 1.67) | 0.001 |

**aCI, confidence interval.**
**bOR, odds ratio.**
**cSD, standard deviation.**
**dBMI, body mass index.**
**eModel 1 adjusted for energy intake (continuous) and age (continuous).**
**fModel 2 adjusted for energy intake (continuous), age (continuous), sex (categorical), education levels (categorical), ethnicity/race (categorical), PIR (continuous).**
**gModel 3 adjusted for energy intake (continuous), age (continuous), sex (categorical), education levels (categorical), ethnicity/race (categorical), PIR (continuous), BMI (categorical), hypertension (categorical), CVD (categorical), cancer (categorical), serum cotinine (continuous), and physical activity (continuous).**
**hP-value for interaction of the Model 3.**
The relationship between DII and prediabetes was substantially unchanged in models with strict criteria to define prediabetes. In addition, a total of 690 participants with undiagnosed diabetes were identified, and we observed a significant relationship between DII and increased odds of undiagnosed diabetes (OR: 1.31; 95% CI: 1.16 to 1.50).

**DISCUSSION**

In this cross-sectional study of American adults, DII was significantly associated with increased levels of FPG, FSI, and HOMA-IR, and was associated with the increased odds of IR and prediabetes after adjustment for potential confounders. Similar findings yielded in analysis stratified by BMI. Besides, restricted cubic spline models showed linear dose–response relationships between DII and the odds of insulin resistant and prediabetes.

The DII score computed from 31 food parameters, ranged from $-5.83$ to $+5.32$ with a mean of $-0.14$, which indicates an anti-inflammatory diet. These results were consistent with prior studies conducted in Mexico, Spain, Italy, and France (35–39)

Notably, for this study, subjects who were older, female, and those with hypertension were more likely to have anti-inflammatory diet, which was similar to a case-control study from Bahrain (40) The exact reason for these observed findings is unclear, we speculated that those adults might be more likely to pay more attention to what they are eating and stay healthy.

The DII score computed from 31 food parameters, ranged from $-5.83$ to $+5.32$ with a mean of $-0.14$, which indicates an anti-inflammatory diet. These results were consistent with prior studies conducted in Mexico, Spain, Italy, and France (35–39)

Notably, for this study, subjects who were older, female, and those with hypertension were more likely to have anti-inflammatory diet, which was similar to a case-control study from Bahrain (40) The exact reason for these observed findings is unclear, we speculated that those adults might be more likely to pay more attention to what they are eating and stay healthy.
only 1,352 participants in the Luxembourg study (19), while the present study included 7,926 participants for primary analysis. In the Iranian study, the participants that were residents of district 13 of Tehran, were not a representative sample of Iranian populations (17). The Brazil study included only young adults aged 23–25 years (41), while the present study included both men and women aged 18 years or older. However, the findings of the present study were in line with the results of a Dutch cohort study, which showed that the DII was positively associated with IR and fasting glucose concentration (16). Additionally, in a more recent cross-sectional study conducted in South Africa, DII was positively associated with fasting glucose, insulin and HOMA-IR (18). Only one study reported the correlation between DII and prediabetes, suggesting that subjects who consumed a pro-inflammatory diet were at increased risk of prediabetes compared with those who consumed a more anti-inflammatory diet (42), whereas, the abovementioned three studies were equally small in sample size. The present research seems to overcome the above mentioned shortcomings and may add new evidence to the limited literature to elucidate the role of DII on the markers of T2D risk, prediabetes, and IR.

Biologic mechanisms underlying the associations of DII with prediabetes and IR remain unclear. Nonetheless, previous studies suggested that pro-inflammatory diet could promote low-grade inflammation (43) that is characterized by elevated pro-inflammatory markers, namely, CRP, TNF-α, and interleukin-6 which may lead to prediabetes and IR (44–46). Furthermore, high intakes of saturated fat and animal protein (47), which are the dietary constituents for DII calculation, may have adverse effects on prediabetes and IR (48). Lastly, studies in the past found that higher DII scores could promote the development of obesity that further contributed to the development of prediabetes and IR (49, 50). However, the present results showed that the associations of DII with IR and prediabetes were significant, with or without adjustment for BMI, indicating the associations of DII with IR and prediabetes might be independent from BMI. In addition, in the stratified models, the present study demonstrated significant associations of DII with IR and prediabetes in participants with or without obesity, and the interaction of DII and BMI was not found to reach statistical significance, implying that the effect of DII on IR and prediabetes might be homogenous in participants with or without obesity. In view of above plausible reasons, future longitudinal cohort and experimental studies should be carried out to reveal the specific causal mechanisms. Further, participants with prediabetes or IR were more likely to develop diabetes and CVD (51, 52), and previous studies have shown that pro-inflammatory diet is associated with increased risk of diabetes and CVD (14, 53). The findings of the present study may further add new evidence to explain their interconnectedness and add new insights into the mechanisms of the occurrence and development of diabetes and CVD.

The present study was based on a large nationally representative survey, which allowed to adjust for multiple covariates and increased the statistical power of the findings. Despite its strengths, we also acknowledge that this study was subject to some limitations. First, the temporal or causal relationships cannot be inferred because of the cross-sectional design for NHANES. However, the study participants were free from known diabetes, so they would not change their diet due to the diagnosis, and thus the dietary information collected at the same time when blood samples were collected could reflect participants’ usual dietary habits. Second, although multicovariates were included in models, we could not exclude the residual confounding effect from unmeasured or unavailable covariates (e.g., genetic factors and use of anti-inflammatory medications). Third, only 31 out of 45 parameters were extracted from 24-hour recalls due to the questionnaire setting. However, previous studies have shown that there was no degradation in predictive ability of the DII in calculation using only 27 or 28 parameters (54, 55). Fourth, although two days of dietary recall were used to reduce the measurement bias, dietary intake of individuals estimated by the average of two 24-hour dietary recalls instead of the NCI method (56) (a useful method to reduce within-person variation and produce a more precise estimate when using NHANES dietary data) does not account for day-to-day variations in diet or seasonal variability in diet patterns. Fifth, subjects included in the present study are more likely to be those apparently healthy adults, which may limit the generalizability of these findings. Lastly, participants with missing data on covariates were excluded, which may limit the generalizability of these findings.

CONCLUSIONS

In conclusion, the present study suggests that DII was positively associated with FPG, FSI, and HOMA-IR, and a more pro-inflammatory diet (expressed as a higher DII score) may increase the odds of insulin resistant and prediabetes. Therefore, the promotion of anti-inflammatory diet might maintain glucose-insulin homeostasis, and subsequently lower the odds of insulin resistant and prediabetes.

DATA AVAILABILITY STATEMENT

The raw data supporting the conclusions of this article will be made available by the authors, without undue reservation.

ETHICS STATEMENT

The studies involving human participants were reviewed and approved by the National Center for Health Statistics research ethics review board. The patients/participants provided their written informed consent to participate in this study.

AUTHOR CONTRIBUTIONS

YS, XW, JW, XM, HL, and YX conceived and designed the research. YS wrote the manuscript draft. YS performed the
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SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fendo.2022.820932/full#supplementary-material
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Conflict of Interest: The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

The reviewer YBZ declared a shared affiliation with several of the authors, YS, XM, YX, to the handling editor at time of review.

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