Dealing with Interference in Distributed Large-scale MIMO Systems: A Statistical Approach

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Abstract—This paper considers the problem of interference control through the use of second-order statistics in massive MIMO multi-cell networks. We consider both the cases of co-located massive arrays and large-scale distributed antenna settings. We are interested in characterizing the low-rankness of users’ channel covariance matrices, as such a property can be exploited towards improved channel estimation (so-called pilot decontamination) as well as interference rejection via spatial filtering. In previous work, it was shown that massive MIMO channel covariance matrices exhibit a useful finite-rank property that can be modeled via the angular spread of multipath at a MIMO uniform linear array. This paper extends this result to more general settings including certain non-uniform arrays, and more surprisingly, to two dimensional distributed large scale arrays. In particular our model exhibits the dependence of the signal subspace’s richness on the scattering radius around the user terminal, through a closed form expression. The applications of the low-rankness covariance property to channel estimation’s denoising and low-complexity interference filtering are highlighted.

Index Terms—massive MIMO, distributed antennas, channel estimation, interference mitigation, covariance matrix.

I. INTRODUCTION

FULL spatial reuse of the frequency resource across even neighboring cells is a de facto standard approach in wireless network design. The downside of this strategy lies in the high amount of inter-cell interference, which in turn severely limits the performance of certain users, especially at cell-edge. This fact has fueled extensive research on interference management, and particularly on methods relying on the use of spatial filtering at the base station side. Recently, two schools of thought have emerged with conflicting strategies for how to best exploit the added spatial dimension offered by multiple-input multiple-output (MIMO) antennas. In the first, the focus is on strengthening local beamforming capabilities by endowing each base station with a massive number of antenna elements that is substantially larger than the number of terminals served in the same cell on any given spectral resource block. The added cost of hardware is compensated by the fact that simple distributed beamforming schemes that require little inter-cell cooperation can efficiently mitigate interference [1]–[4]. In the second school, cooperation between cells is emphasized as the key towards increasing the spatial degrees of freedom [5]. In the cooperation approach, so-called network MIMO (or CoMP in the 3GPP terminology) schemes mimic the transmission over a virtual MIMO array encompassing the spatially distributed base station antennas. In contrast with the massive MIMO solution, the cooperative spatial filtering of interference is made possible with no additional antennas at the base station side, yet it goes at the expense of fast signaling links over the backhaul, a need for tight synchronization, and seemingly multi-user detection schemes that are computationally more demanding than the simple matched filters advocated in massive MIMO. Additionally, a major hurdle preventing from realizing the full gains of MIMO multi-cell cooperation lies in the cost of acquiring and sharing channel estimates using orthogonal training sequences over large clusters [6].

Despite these differences, a fundamental common feature behind each philosophy lies in the coherent combining of a large number of antennas in view of interference nulling. Additionally, in both cases, our ability to reject interference is only as good as our ability to estimate the user channels properly. In the context of co-located massive MIMO, channel estimation from pilots that are inevitably reused over space leads to the so-called pilot contamination effect [7], [8]. Although initially branded as a fundamental limit of massive MIMO communications, a finer impact analysis of pilot contamination indicates that it is only one of several limitations of such systems [9]. When it comes to improving channel estimation, several possible solutions were recently proposed in a series of papers [10]–[12]. In [10], an approach to de-interfere channel estimates was revealed based on the exploitation of second-order statistical properties of the received vector signal. The key enabler is the finite-rankness of the channels’ covariance matrices which was shown to occur in the asymptotic massive MIMO regime whenever the angle spread of incoming/departing paths at the MIMO array is limited. Independently, a similar finite-rank property was shown to be useful in the context of low-complexity scheduling and spatial beamforming for massive MIMO networks [13]. Hence the low-dimensional property for the signal subspace (i.e. in which the MIMO channel realizations live) is instrumental to spatial interference rejection. These results were all reached for the case of uniform (equi-spaced) calibrated linear arrays. A natural question then arises as to whether the low-rank property can be established and exploited in more general large-scale antenna settings, such as random and two dimensional antenna placements. This paper is devoted to this problem.
A first examination of [10], [13] indicates that the finite-rank behavior is rooted in the asymptotic orthogonality between Fourier transform vectors corresponding to different path angles, suggesting the property might be restricted to the use of one-dimensional equi-spaced arrays. However our results point otherwise, showing low-rankness of channel’s subspace system is a recurrent trend applying to random and also distributed antenna placements, hence with a wider applicability to cooperative networks.

Our specific contributions are as follows: First we consider a uniform linear massive array scenario yet with several clusters of multipath. In this case we establish a finite-rank model for the channel’s covariance that directly extends that of [10], where the rank is shown to be a function of the incoming/departing angular spread of multipath. We then show that a similar low-rank result holds for a linear array with random placement of antenna elements. Although in this case, unlike the uniform array, the finite rank is only characterized by an upper bound. We show how this property can be used towards, for instance, pilot decontamination.

In the second part of the paper, we turn to a large-scale antenna regime where the antenna elements are scattered randomly throughout the (dense) network, yet can still be combined coherently. Such a setting with spatially distributed antennas includes remote radio head (RRH) networks, network-MIMO (CoMP) schemes with large clusters, and cloud-enabled radio access networks (C-RAN) as particular cases. A channel model building on the classical one-ring multipath model [14], [15] is proposed to analyze this scenario. In this setting we show that, there again surprisingly, the channel covariance exhibits a low-dimensional signal subspace behavior, in the large number of base station antenna regime, even discounting path loss effects. We show the richness of the covariance’s signal subspace is primarily governed by the scattering radius around the user terminal. We provide a closed form expression for an upper-bound of the covariance rank and show by simulation how this bound closely matches reality. Note that the notion that the total perimeter occupied by scatterers can govern the rank of the signal subspace in a distributed MIMO antenna setting is reminiscent of a previously observed phenomenon in the different context of compact MIMO arrays. In [16], the authors establish a physical model for the dimension of the spatial multipath field of a disk-shaped compact area filled with MIMO antennas and illuminated by isotropic multipaths.

In the last part of the paper, we turn our attention to the exploitation of signal-subspace’s low-rankness towards interference rejection for a distributed array. We derive a lower bound on the signal to interference ratio that would be obtained in a two user setting with a simple matched filter, as a function of the distance between the users and the number of antennas. We show how a distance of two scattering radii can be selected as a critical minimal distance between selected co-channel users in a scheduling algorithm so as to facilitate interference nulling. As an application of the low-rankness property, a simple subspace-based interference mitigation scheme is put forward, which exploits the statistical information of the interference channels. Numerical results are presented in the last section.

The notations adopted in the paper are as follows. We use boldface to denote matrices and vectors. Specifically, $I_M$ denotes the $M \times M$ identity matrix. $(X)^T$, $(X)^*$, and $(X)^H$ denote the transpose, conjugate, and conjugate transpose of a matrix $X$ respectively. $E\{\cdot\}$ denotes the expectation, $\|\cdot\|_F$ denotes the Frobenius norm. The Kronecker product of two matrices $X$ and $Y$ is denoted by $X \otimes Y$. $\text{span}\{v_1, v_2, \ldots, v_n\}$ is the span of linear vector space on the basis of $v_1, v_2, \ldots, v_n$ for some $n \geq 1$. $\dim\{A\}$ is the dimension of a linear space $A$, and $\text{null}\{R\}$ is the null space of matrix $R$. $\text{diag}\{a_1, \ldots, a_N\}$ denotes a diagonal matrix or a block diagonal matrix with $a_1, \ldots, a_N$ at the main diagonal. $\Xi$ is used for definition.

II. Co-located Massive Linear Arrays

We consider the uplink\footnote{Similar principles would apply in the downlink, which for ease of exposition is ignored here.} of a network of $B$ time-synchronized cells, with full spectrum reuse. Each of the $B$ base stations is equipped with a one-dimensional array of $M$ antennas, where $M$ is allowed to grow large (massive MIMO regime). For ease of exposition, all user terminals are assumed to be equipped with a single antenna. Furthermore we consider that a single user is served per cell and per resource block. A classical multipath model is given by [17]:

$$h_i = \sqrt{\frac{\beta_i}{P}} \sum_{p=1}^{P} a(\theta_{ip}) e^{j\phi_{ip}},$$

(1)

where $P$ is the arbitrary number of i.i.d. paths, $\beta_i$ denotes the path loss for channel $h_i$, and $e^{j\phi_{ip}}$ is the i.i.d. random phase, which is independent over channel index $i$ and path index $p$. $a(\theta)$ is the signature (or phase response) vector by the array to a path originating from the angle $\theta$. Note that in the case of an equi-spaced array, $a(\theta)$ has a Fourier structure.

A. Channel Estimation

When it comes to channel estimation it is assumed that orthogonal pilots are used by users located in the same cell, so that intra-cell pilot interference can be neglected. Sets of pilot sequences are however assumed to be fully reused from cell to cell, causing maximum inter-cell pilot interference. The pilot sequence is denoted by:

$$s = [s_1 \ s_2 \ \cdots \ s_r]^T.$$

(2)

The power of the pilot sequence is assumed to be $s^H s = \tau$. The channel vector between the $b$-th cell user and the target base station is $h_b$. Without loss of generality, we assume the 1st cell is the target cell. Thus, $h_1$ is the desired channel while $h_b, b > 1$ are interference channels. During the pilot phase, the signal received at the target base station is

$$Y = \sum_{b=1}^{B} h_b s^T + N,$$

(3)

where $N \in \mathbb{C}^{M \times \tau}$ is the spatially and temporally white additive Gaussian noise (AWGN) with zero-mean and element-wise variance $\sigma_n^2$. Assuming the desired and interference
covariance matrices $\mathbf{R}_b \triangleq \mathbb{E}\{\mathbf{h}_b\mathbf{h}_b^H\}$ can be estimated in a preamble, the Bayesian (or equivalently MMSE) estimate of the target channel vector is given by [3, 10, 18]:

$$\hat{\mathbf{h}}_1 = \mathbf{R}_1 \left( \sigma_n^2 \mathbf{I}_M + \tau \sum_{b=1}^{B} \mathbf{R}_b \right)^{-1} \bar{\mathbf{S}}^H \mathbf{y}, \quad (4)$$

where the training matrix $\bar{\mathbf{S}} \triangleq \mathbf{s} \otimes \mathbf{I}_M$ and $\mathbf{y} \triangleq \text{vec}(\mathbf{Y})$. An interesting question is under which conditions $\hat{\mathbf{h}}_1 \rightarrow \hat{\mathbf{h}}_1^{\text{no int}}$ in the massive MIMO regime ($M \gg 1$), where the superscript "no int" refers to the "no interference case." This question was previously addressed in [10], revealing the following sufficient condition for achieving total interference suppression in the large $M$ regime:

$$\bigcup_{b=2}^{B} \text{span} \left\{ \mathbf{R}_b \right\} \subset \text{null} \left\{ \mathbf{R}_1 \right\} \quad (5)$$

where the above condition requires the target channel covariance to exhibit a non-empty null space (aka low-dimensional subspace) and for all other interference covariances’ signal subspaces to fall within this null space (see the proof in [10]).

In practice, the inclusion condition in (5) can be realized by a user grouping algorithm [10], [13], as long as the rank of each covariance is small enough in relation to $M$.

B. Low-rank Properties of General Linear Arrays

In [10] [13], a linear equi-spaced array was considered. The propagation model also assumed that multipaths impinge on the base station array with angles of arrival (AOA) spanning an interval $[\hat{\theta}_q^{\text{min}}, \hat{\theta}_q^{\text{max}}] \in [0, \pi]$ [12]. It is then shown that condition (5) is satisfied provided AOAs corresponding to interfering users fall outside $[\hat{\theta}_q^{\text{min}}, \hat{\theta}_q^{\text{max}}]$. The assumptions of a single cluster of multipath and of a calibrated equi-spaced array are however restrictive. Below, we generalize this result to more realistic settings.

1) Multiple Clusters: We now consider a general multipath model when the AOAs corresponding to the desired channel are still bounded, but come from several disjoint clusters [17]. The steering vector in (1) is [19]

$$\mathbf{a}(\theta) \triangleq \begin{bmatrix} 1 \\ e^{-j2\pi \frac{\lambda}{D} \cos(\theta)} \\ \vdots \\ e^{-j2\pi \frac{(M-1)D}{\lambda} \cos(\theta)} \end{bmatrix}, \quad (6)$$

where $D$ is the antenna spacing and $\lambda$ is the signal wavelength. Let $Q$ denote the number of clusters. Let $[\hat{\theta}_q^{\text{min}}, \hat{\theta}_q^{\text{max}}]$ denote the interval of AOAs for the $q$-th cluster of desired paths in the $[0, \pi]$ interval. See an illustration in Fig. 1 for $Q = 2$.

For a uniform linear array, we have the following proposition in the massive MIMO regime:

**Proposition 1.** The rank of channel covariance matrix $\mathbf{R}$ satisfies:

$$\frac{\text{rank}(\mathbf{R})}{M} \leq d, \quad \text{when } M \text{ is sufficiently large},$$

$^2$Note that a path coming from angle $-\theta$ yields identical steering vector to that from $\theta$. Therefore we can limit ourselves to AOAs within $[0, \pi]$.

where $d$ is defined as

$$d \triangleq \min(1, \sum_{q=1}^{Q} (\cos(\hat{\theta}_q^{\text{min}}) - \cos(\hat{\theta}_q^{\text{max}})) \frac{D}{\lambda}).$$

**Proof:** The channel can be seen as the sum of elementary channels each of which corresponds to one separate clusters. Then $\mathbf{R}$ can be decomposed into a sum of covariances over these clusters. Since the clusters are separated, the signal subspaces of the corresponding covariances are orthogonal and therefore their dimensions add up. Then based on [10] Lemma 1, the proof of Proposition 1 can be readily obtained.

Now define the total set of AOAs of the desired channel as

$$\bar{\Theta}_d \triangleq \bigcup_{q=1}^{Q} [\hat{\theta}_q^{\text{min}}, \hat{\theta}_q^{\text{max}}],$$

so that the probability density function (PDF) $p_d(\theta)$ of the desired AOA satisfies $p_d(\theta) > 0$ if $\theta \in \bar{\Theta}_d$ and $p_d(\theta) = 0$ if $\theta \notin \bar{\Theta}_d$. In the same way, the PDF of all interference AOAs satisfies $p_i(\theta) > 0$ if $\theta \in \bar{\Theta}_i$ and $p_i(\theta) = 0$ otherwise, where $\bar{\Theta}_i$ is the union of all possible interference AOAs. We have the following result for the massive uniform array:

**Corollary 1.** If $D \leq \lambda/2$ and $\bar{\Theta}_d \cap \bar{\Theta}_i = \emptyset$, then the MMSE estimate of (4) satisfies:

$$\lim_{M \to \infty} \hat{\mathbf{h}}_1 = \hat{\mathbf{h}}_1^{\text{no int}}.$$  

**Proof:** It can be shown that from [10] Lemma 2, condition (5) will be fulfilled as long as interfering AOAs do not overlap with any of the clusters for the desired channel, in which case if we analyze the received signal using eigenvalue decomposition, we can find the interference disappears asymptotically because of its orthogonality with the signal space of desired channel covariance. (8) is obtained in the same way as [10]. As a result we omit the detailed proof in this paper.

2) Random Arrays: Tightly calibrated arrays with uniform spacing are hard to realize in practice. An interesting question is whether the above results carry on to the setting of linear arrays with random antenna placement. To study this case, we consider a set of antennas randomly located over a line, and spanning a total aperture of $D$ meters. We investigate the extended array and $D$ is allowed to grow with $M$.  

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**Fig. 1.** Desired channel composed of $Q = 2$ clusters of multipath.
In this case, an elementary path coming from an angle \( \theta \) can be represented via the corresponding array response vector as:

\[
\mathbf{a}(\theta) \triangleq \left[ e^{-j2\pi \frac{d_1}{\lambda} \cos(\theta)}, \ldots, e^{-j2\pi \frac{d_M}{\lambda} \cos(\theta)} \right]^T,
\]

where the position of the \( m \)-th antenna \( \bar{D}_m \) (\( 1 \leq m \leq M \)), \( d_m \), follows a uniform distribution, i.e., \( d_m \sim \mathcal{U}(0, \bar{D}) \). The PDF of AOA \( \theta \) for the desired paths is non-zero only when \( \theta \in \bar{B}_D \), as in section II-B. Define the average antenna array \( \bar{D} \) as in (7), follows a uniform distribution, i.e., \( \bar{D} \sim \mathcal{U}(0, \bar{D}) \). The following result now directly generalizes Proposition 1 to massive MIMO regime, i.e. for \( M \to +\infty \). In the following proposition we extend the results of Corollary 1 to the case of random arrays under the weaker assumption of rank additivity for the covariance matrices of the desired and interference channels.

**Proposition 3.** With a bounded support of AOAs \( \bar{B}_d \) as in (7), the rank of channel covariance matrix \( \mathbf{R} \) satisfies:

\[
\text{rank}(\mathbf{R}) \leq \sum_{q=1}^{Q} (\cos(\theta_q^\text{min}) - \cos(\theta_q^\text{max})) \frac{M\bar{D}}{\lambda} + o(M),
\]

**Proof:** We can readily obtain this result by replacing \( x \) with \( \cos(\theta) \) in Proposition 2. This result above suggests that the low-dimensional feature of signal subspaces in massive MIMO is not critically linked to the Fourier structure of the steering vectors. Furthermore, it should be noted that the above upper bound is actually very tight for large \( M \), as witnessed from the simulation in Fig. 2 where we take \( Q = 1, \bar{D} = \lambda/2 \) for example. The AOA spread is 40 degrees, and the closed form model refers to

\[
\sum_{q=1}^{Q} (\cos(\theta_q^\text{min}) - \cos(\theta_q^\text{max})) \frac{M\bar{D}}{\lambda}.
\]

We can observe that \( \text{rank}(\mathbf{R}) \) is well approximated by \( f(M) \).

**Note:** Note that antenna ordering has no impact on our results.

![Image](image-url)

Fig. 2. Closed-form rank model for the channel covariance vs. actual rank.
Let us observe that
\[ T^\dagger R_d = 0 \quad \text{and} \quad R_d S^\dagger = 0. \quad (13) \]
We focus on the first equality. The proof of the second equation follows along the same line. By appealing to the mixed type reverse order laws of the \( r \times s \) matrix \( A \) and the \( s \times t \) matrix \( B \) in \([22]\)
\[ (AB)^\dagger = B^H (A^H ABB^H)^\dagger A^H, \]
\( T^\dagger \) can be rewritten as
\[ T^\dagger = \bar{U}_i \bar{U}_i^H \left[ (I - \bar{U}_d \bar{U}_d^H) \bar{U}_i \bar{U}_i^H \right]^\dagger (I - \bar{U}_d \bar{U}_d^H) \]
\[ = \bar{U}_i \bar{U}_i^H T^\dagger (I - \bar{U}_d \bar{U}_d^H). \]
The first equality is obtained utilizing the fact that the matrices \( \bar{U}_i \bar{U}_i^H \) and \( (I - \bar{U}_d \bar{U}_d^H) \) are orthogonal projectors and thus idempotent. Then,
\[ T^\dagger R_d = \bar{U}_i \bar{U}_i^H T^\dagger (I - \bar{U}_d \bar{U}_d^H) \bar{U}_d \Sigma_d \bar{U}_d^H = 0. \]
Finally, substituting (12) into (11) and accounting for orthogonality in (13)
\[ C_e = R_d - \bar{R}_d \left[ (I - S^\dagger R_d^\dagger (I - T^\dagger) + S^\dagger R_d^\dagger T^\dagger) \right] R_d \]
\[ = R_d - \bar{R}_d \bar{R}_d^\dagger R_d = 0. \]
In the last equality we use one of the fundamental relations defining the Moore-Penrose generalized inverse.

According to Proposition 3, the rank additivity condition is in general satisfied when the AOA support of desired channel and that of interference channels span disjoint region of spaces, i.e., \( \bar{\theta}_d \cap \bar{\theta}_i = \emptyset \). This property can be exploited in pilot decontamination or interference rejection. Fig. 3 shows the channel estimation performance in the presence of contaminating pilots. In the simulation, we consider a 2-cell network. Each cell has one single-antenna user who uses identical pilot sequence. The mean squared error (MSE) of uplink channel estimation is shown. The simulation suggests that the MMSE channel estimator is able to rid itself from pilot contamination effects as the number of antennas is (even moderately) large, which verifies Proposition 4.

**III. Finite Rank Model in Distributed Arrays**

We now turn to another popular form of large scale antenna regime, often referred to in the literature as distributed antenna systems. In such a setting, a virtual base station is deployed having its \( M \) antennas scattered throughout the cell\(^5\). We consider again the uplink in which joint combining across all BS antennas is assumed possible. The \( M \) base station antennas are assumed uniformly and randomly located in a fixed size network, serving single-antenna users. \( M \) is allowed to grow large giving rise to a so-called dense network. Our model assumes a disk-shaped cell of radius \( L \), although simulation and intuition confirm that the actual shape of the cell’s boundary is irrelevant to the main result.

\(^5\)For ease of exposition we temporarily consider a single cell setting in this section, i.e., \( B = 1 \). However simulation is also done later in a multi-cell scenario.

\[ \beta_{kpm} = \frac{\alpha}{(d_{kpm} + r)^\gamma}, \quad (14) \]
where \( \alpha \) is a constant that can be computed based on desired cell-edge SNR, and \( \gamma \) is the path loss exponent. We scale the amplitude of each path by \( \sqrt{P} \). The channel between user \( k \) and all BS antennas is given by:
\[ h_k \triangleq \frac{1}{\sqrt{P}} \sum_{p=1}^{P} h_{kp}, \quad (15) \]
where \( h_{kp} \) is the \( p \)-th scattering path vector channel between user \( k \) and all base station antennas:
\[ h_{kp} \triangleq \left[ \begin{array}{c}
\sqrt{\beta_{kp1} e^{-i2\pi \frac{d_{kp1}+r}{\lambda}}} \\
\vdots \\
\sqrt{\beta_{kpM} e^{-i2\pi \frac{d_{kpM}+r}{\lambda}}} \\
\end{array} \right] e^{j\varphi_{kp}}, \quad (16) \]
where \( e^{j\varphi_{kp}} \) denotes the random common phase of that scattering path vector due to possible random perturbations of the user location around the ring center or the phase shift due to the reflection on the scatterer. \( \varphi_{kp} \) is assumed i.i.d. and uniformly distributed between 0 to \( 2\pi \).

\(^6\)Note that this model assumes the BS antennas are high enough above clutter so that there is no local scattering around the BS antennas.
B. A Low-rank Model for Distributed Arrays

We have shown in section II-B the low-dimension property for linear antenna array systems. In attacking this problem it is important to distinguish the rank reduction effect due to path loss from the intrinsic finite-rank behavior of the large antenna channel covariance in an equal path loss regime. In fact, in an extended network (i.e. where some base station antennas can be arbitrarily far from some users), the signal of any given user will be received over only a limited number of antennas in its vicinity, thereby effectively limiting the channel rank to the size of this neighborhood. To circumvent this problem, we consider below a (dense) network where the path loss terms are set artificially to be all equal (to one) and study the finite-rankness under such conditions. In this model, the channel covariance is defined as $R \triangleq \mathbb{E}\{hh^H\}$ where the expectation is taken over the random positions of the scatterers on the ring. Note that our analysis indicates that a randomization over the user’s location inside the scattering’s disk would produce an identical upper bound on the rank.

**Theorem 1.** The rank of the channel covariance matrix for a distributed antenna system satisfies:

$$\text{rank}(R) \leq \frac{4\pi r}{\lambda} + o(r).$$

**Proof:** See Appendix B.

In reality we show below that the right hand side of (17) is a very close approximation of the actual rank, which is defined as the number of eigenvalues of $R$ which are greater than a prescribed threshold (in our simulations it is taken to be 10e-5). Theorem 1 shows a linear dependency of the rank as the number of eigenvalues of $R$ which are greater than a prescribed threshold (in our simulations it is taken to be 10e-5). Theorem 1 shows a linear dependency of the rank.

**Proposition 5.** For $D_{pq}$ small enough that the two scatterers are located in the same planar wavefront region, we have

$$\lim_{M \to \infty} \frac{|h_{2q}|^2 |h_{1p}|}{|h_{1p}|^2 |h_{2q}|} \approx J_0 \left( \frac{2\pi D_{pq}}{\lambda} \right),$$

where $J_0$ is the zero-order Bessel function of the first kind.

**Proof:** See Appendix C.

Note that the proof is based on an additional assumption that the path loss between a certain antenna and the two scatterers are approximately equal, i.e., if user 1 and user 2 are concerned, then $\beta_{2qm} \approx \beta_{1pm}$. Since the two scatterers are very close and the antenna is much further away, this
5. We draw the curves of \(|h_{2q}^H h_{1p}|\) and \(|J_0(2\pi D_{pq}/\lambda)|\) in Fig. 6. The curves show that when \(M\) grows, \(|h_{2q}^H h_{1p}|/|h_{1p}|\) gets closer and closer to the Bessel function. The curve named “Envelope by Krasikov” is an upper bound of the Bessel function developed by Krasikov in [23], which will be used in Proposition 6. The interpretation of Proposition 5 is as follows: the highest correlation between two scattering paths is attained when the spacing between the two scatterers is very small. Since the Bessel function reaches its first zero around a spacing of \(D_0 = 0.38\lambda\), it implies that the users ought to be located at least \(2r + D_0 \approx 2r\) away from each other to allow for a reuse of spectral resources (such as pilots). In practice we would intend to schedule users with additional spacing than just \(2r\), according to the side lobes of the Bessel function. Note that our model is similar to Clarke’s model [24], which holds, making the use of the Bessel function impractical.

We use Proposition 5 to derive a lower bound on the signal to interference ratio (SIR) under a simplified system setting with just two users (one desired, one interferer) and a matched filter receiver [20]. By computing the expectation over the random BS antenna locations, which is also how we derive the expectation of SIRs and channel correlations in the rest of this section, we obtain the following bound of SIR:

### Proposition 6

**Assume perfect channel estimation, closely located users, whose scattering rings do not overlap. The expected SIR at the matched filter output satisfies:**

\[
\mathbb{E}\{\text{SIR}\} \geq \frac{\pi \left( \left( 4\left(\frac{2\pi(D_u - 2r)}{\lambda}\right)^2 - 3 \right) \frac{1}{2} - 3 \right)}{16\left(\frac{2\pi(D_u - 2r)}{\lambda}\right)^2 - 20}, \text{ when } M \text{ is large,}
\]

where \(D_u\) is the distance between the two users and is assumed to be larger than \(\sqrt{3 + 3^4 \lambda/(4\pi) + 2r}\).

\[\text{Proof:}\] When applying matched filter, we have

\[
\mathbb{E}\{\text{SIR}\} = \mathbb{E}\left[ \frac{|h_1|^2 |h_2|^2}{|h_2^H h_1|^2} \right].
\]

Let us recall that the envelope of the Bessel function \(J_0\left(\frac{2\pi D}{\lambda}\right)\) is decreasing with \(D\). Thus, the lower bound of SIR is obtained by considering the shortest \(D_{pq}\), which gives the worst case correlation. Since \(D_{pq} \geq D_u - 2r\), we may obtain when \(M\) is large:

\[
\mathbb{E}\{\text{SIR}\} \geq \frac{1}{\left|J_0\left(\frac{2\pi(D_u - 2r)}{\lambda}\right)\right|^2}.
\]

Finally we use an upper bound of the envelope of the Bessel function [23] which has validity when \(D_u - 2r > \sqrt{3 + 3^4 \lambda/(4\pi)}\). The bounding argument of the Bessel function in [23] can directly apply here.

The above proposition quantifies the rate at which the SIR increases with the inter-user distance, in this case linearly.

2) **Distant Users:** We consider the regime in which users are located further away from each other, e.g., many wavelengths away. The planar wavefront assumption no longer holds, making the use of the Bessel function impractical.

In this case we are again interested in characterizing the correlation between two scattering paths corresponding to two users, then the correlation between the channel vectors themselves.

We first investigate the behavior of \(h_{2q}^H h_{1p}\) for any \(p, q\):

\[
h_{2q}^H h_{1p} = \sum_{m=1}^{M} (h_{2qm}^* h_{1pm}),
\]

where \(h_{kpm}\) is the channel between the \(k\)-th user and the \(m\)-th BS antenna via the \(p\)-th scatterer:

\[
h_{kpm} = \sqrt{\beta_{kpm}} e^{-j2\pi \frac{d_{kpm}}{\lambda} + j\varphi_{kpm}}.
\]

Since the two phases \(e^{j\varphi_{1p}}\) and \(e^{j\varphi_{2q}}\) are independent, we have:

\[
\mathbb{E}(h_{2qm}^* h_{1pm}) = 0,
\]

and the variance of \(h_{2qm}^* h_{1pm}\) is

\[
\sigma^2(D_{pq}) \triangleq \text{Var}(h_{2qm}^* h_{1pm}) = \mathbb{E}(\beta_{2qm} \beta_{1pm}).
\]

Given the random network model with radius \(L\), the path loss correlation can be found by integration over polar coordinates giving the location of the BS antennas. Although a closed-form expression is elusive, we get the following computable expression:

\[
\sigma^2(D) = \frac{2\nu}{\pi L^2} \int_0^L dp \int_0^\pi \rho \rho \left( \sqrt{D^2 + \rho^2 - 2\rho D \cos(\varphi)} + r \right) \sin(\varphi) d\varphi.
\]

One example of the path loss correlation \(\sigma^2(D)\) is given in Fig. 7 which shows \(\sigma^2(D)\) is a decreasing function of \(D\). We have the following proposition on the distribution of \(|h_{2q}^H h_{1p}|^2|:\n
\[
\text{with more users, the interference is simply scaled by the number of users. Additionally more advanced receivers could also be exploited.}
\]
where $h$ denotes convergence in distribution.

Proof: Let us consider the random variable

$$S_M = \frac{\mathbf{h}_{2q}^H \mathbf{h}_{1p}}{\sigma^2(D_{pq})M}.$$ 

By the law of large numbers, $S_M$ converges almost surely to 0 as $M \to +\infty$. By appealing to the Central Limit Theorem (CLT), the random variable $\sqrt{M}S_M$ converges in distribution to a complex Gaussian distribution with zero mean and unit variance. Thus, its square, i.e., $M|S_M|^2$, converges also in distribution to an exponentially distributed random variable with parameter $\lambda = 2$.

We now derive a lower bound of the average SIR for a two-user system, assuming matched filtering receiver.

**Proposition 8.** The lower bound is given by

$$\mathbb{E}\{\text{SIR}\} \geq \frac{M(C^2 + o(1))}{\sigma^2(D_u - 2r)},$$

where $C$ is a constant such that $h_1^H h_1 / M = C + o(1)$.

Proof: The SIR can be written as:

$$\text{SIR} = \frac{\|h_1^H h_1\|^2}{\|h_2^H h_2\|^2} = \frac{C^2 + o(1)}{M^2}.$$

Note that $\forall p, q$, $h_2^H h_{1p} / M$ has zero mean and a variance of $\sigma^2(D_{pq}) / M$. In addition, the two variables $h_2^H h_{1p} / M$ and $h_2^H h_{1p'} / M$ are uncorrelated for any $p' \neq p$ or $q' \neq q$, resulting from the random and independent phases in (16).

$$\mathbb{E}\left\{\frac{|h_2^H h_1|^2}{M^2}\right\} = \mathbb{E}\left\{\frac{1}{P^2} \sum_{q=1}^{P} \sum_{p=1}^{P} \frac{|h_2^H h_{1p}|^2}{M}\right\} = \frac{1}{P^2} \sum_{q=1}^{P} \sum_{p=1}^{P} \text{Var}\left(\frac{|h_2^H h_{1p}|}{M}\right) \leq \frac{\sigma^2(D_{pq})}{M} \leq \frac{\sigma^2(D_u - 2r)}{M}.$$

The final step is due to the fact that $D_{pq} \geq D_u - 2r, \forall p, q$. Finally we get

$$\mathbb{E}\{\text{SIR}\} \geq \frac{M(C^2 + o(1))}{\sigma^2(D_u - 2r)},$$

and Proposition 8 is proven.

The above result suggests at which rate the matched filtered interference decays as a function of $M$ and the inter-user spacing $D_u$, which in turn can be exploited to predict system performance and also give insights to the tolerable spatial reuse of pilot resources.

### B. Interference Filtering via Subspace Projection

SIR analysis in previous sections is built upon a simple matched filter, which still requires an accurate channel estimation. In this section, however, we propose a simple beamforming strategy building on the low-dimensionality of the signal subspace, which does not require an accurate channel estimation. We consider a $K$-user network with the first user being a target user and all other users being interference users. All these users share the same pilot sequence $s$. Denote the sum of interference covariances as $R_I = R_2 + R_3 + \cdots + R_K$. The eigenvalue decomposition of $R_I$ is $R_I = \Sigma \Sigma^H$, where $\Sigma$ is a $M \times M$ diagonal matrix with the eigenvalues of $R_I$ on its main diagonal. Suppose the eigenvalues are in descending order and the first $m$ eigenvalues are non-negligible while the others can be neglected. We construct the spatial filter at the BS side for user $1$ as:

$$W_1 = [u_{m+1} | u_{m+2} | \cdots | u_M]^H,$$

where $u_m$ is the $m$-th column of $U$. We can assume approximately that:

$$W_1^H h_k \approx 0, \forall k \neq 1,$$

$$W_1^H Y \approx W_1^H h_1 s^T + W_1^H N,$$

where $N \in \mathbb{C}^{M \times \tau}$ is the spatially and temporally white additive Gaussian noise, $Y \in \mathbb{C}^{M \times \tau}$ is the received training signal, and $s \in \mathbb{C}^{\tau \times 1}$ is the shared pilot sequence. Define the effective channel $\hat{h}_1 = W_1^H h_1$. Note that $\hat{h}_1$ has a reduced size, which is $(M - m) \times 1$. An LS estimate of $\hat{h}_1$ is:

$$\hat{h}_1 = W_1 Y s^*(s^T s)^{-1}.$$
The key idea is that channel estimate $\hat{h}$ is coarse, yet it can be used as a modified MRC beamformer as it lies in a subspace orthogonal to the interference and is also aligned with the signal subspace of $h$. During uplink data transmission phase:

$$y = h_1s_1^T + \sum_{k=2}^{K} h_k s_k^T + n,$$

where $s_1, s_2, \cdots, s_K \in \mathbb{C}^{r_u \times 1}$ are the transmitted signal sequence, $y, n \in \mathbb{C}^{M \times r_u}$ are the received signal and noise respectively. The subspace-based MRC beamformer is $\hat{h}_1^H W_1$:

$$\hat{h}_1^H W_1 y = \hat{h}_1^H h_1 s_1^T + \hat{h}_1^H W_1 \sum_{k=2}^{K} h_k s_k^T + \hat{h}_1^H W_1 n. \quad (30)$$

In case there is no null space for $R_f$, e.g., the number of users $K$ is large or the interference users have rich scattering environments, the subspace-based method can still avoid the strong eigen modes of interference and therefore reject a good amount of interference.

Note that the subspace projection method has a certain similarity with [11] which also uses eigen-value decomposition in order to perform blind channel estimation. However there are two main differences: 1) They address only the case of classical massive arrays, not distributed antenna arrays; 2) They use the received power levels domain to separate desired channel and interfering channels. In our approach, the discrimination against interference is related to the phases with which the interference and desired signals arrive at the array. In fact, the two techniques could in principle be combined.

V. NUMERICAL RESULTS

We first consider the channel estimation quality in a random network with radius $L = 500$ meters. The path loss exponent $\gamma = 2.5$. The scattering radius is $r = 15$ meters. $P = 50$ scatterers are randomly distributed in the scattering ring, which is centered at the user. Define the channel estimation MSE of the $k$-th user as:

$$\text{MSE}_k \triangleq 10 \log_{10} \left( \frac{\|\hat{h}_k - h_k\|^2}{\|h_k\|^2} \right). \quad (31)$$

In the simulation we average the channel estimation MSE over different users in order to obtain MSE curve.

In Fig. 8 we assume the target user is located at the origin while an interfering user (they share the same pilot sequence) is moving over the horizontal axis at increasing distances from user 1. As we can observe, when the MMSE estimator (4) is used, the channel estimation error is a monotonous decreasing function of the distance between the desired user and the interference user. One may also notice the constant performance gap between LS and MMSE estimator in interference-free scenario, which indicates that covariance information is still helpful even in a highly distributed antenna system. As shown in the blue curve on the top, an LS estimator is unable to separate the desired channel and the interference channel. In contrast, an MMSE estimator has much better performance as its MSE is decreasing almost linearly with inter-user spacing, hence confirming our claims.

We then examine the performance of four MRC beamformers in terms of uplink sum-rate in a single-cell setting (Fig. 9) and per-cell rate in a multi-cell setting (Fig. 10). The sum-rate is defined as follows:

$$\text{sum-rate} \triangleq \sum_{k=1}^{K} \log_2(1 + \text{SINR}_k), \quad (32)$$

where $K$ is the number of simultaneously served users, and $\text{SINR}_k$ is the uplink signal-to-noise-plus-interference ratio (SINR) of the $k$-th user.

In Fig. 9 we show the performance of subspace-based MRC beamforming in a single-cell network where two users share the same pilot. The total number of distributed antennas is 500. In the figure “LS + MRC” denotes the sum-rate performance of MRC beamforming using the LS channel estimate, while the curve “MMSE + MRC” is the performance of MRC beamforming using the MMSE estimate (4). “MMSE + MMSE” denotes the performance curve of MMSE beamforming using MMSE channel estimate when channel covariances (including the interference covariances) are assumed known during both channel estimation and signal detection. The simulation shows the simple subspace-based method has a very good performance. Due to pilot contamination, the MRC beamformer using MMSE channel estimate is not as good as subspace-based method. The reason is that $R_1$ and $R_2$ generally have overlapping signal subspaces here. We may also notice that the subspace-based MRC beamformer has some slight performance gains over the MMSE beamformer.

Fig. 10 depicts the uplink per-cell rate achieved by the above-mentioned MRC beamformers as a function of scattering radius $r$. In the simulation we have 7 hexagonal cells with one center cell and 6 surrounding cells. Each cell has one user. All the users share the same pilot sequence. The per-cell rate is defined as the sum-rate (32) divided by the number of cells. As can be seen, the subspace-based beamforming
the radius of the scattering ring is larger. It also shows more robustness than MMSE beamformer when especially when the radius of scattering ring is smaller. It shows performance gains over other traditional MRC methods especially when the radius of the scattering ring is smaller.

This shows performance gains over other traditional MRC methods especially when the radius of scattering ring is smaller. It also shows more robustness than MMSE beamformer when the radius of the scattering ring is larger.

**VI. CONCLUSIONS**

We investigate low-dimensional properties of covariance signal subspaces in general topologies of massive arrays. We extend previous results known in the uniform linear array case to the case of arrays with random placement, including the case of scattered antennas over a 2D dense network. A correlation model is derived which is exploited to gain insight on the interference rejection capability of low-complexity matched filter-based receivers in distributed antenna settings.

**APPENDIX**

**A. Proof of Proposition 2**

We first consider an $N$-antenna ULA with aperture $D$ and antenna spacing $D = D/(N - 1)$. Define

$$\tilde{\alpha}(x) \triangleq [0, e^{-j2\pi x}, \ldots, e^{-j2\pi D(N - 1)x}]^T = [0, e^{-j2\pi D/Nx}, \ldots, e^{-j2\pi D(N-1)x}]^T.$$  

Now we define $\tilde{A} \triangleq \text{span}\{\tilde{\alpha}(x), x \in [b_1, b_2]\}$. Recall from [10] the following result: If $\beta(x) \triangleq [1, e^{-j\pi x}, \ldots, e^{-j\pi(N-1)x}]^T$. Given $b_1, b_2 \in [-1, 1]$ and $b_1 < b_2$, define $A \triangleq \text{span}\{\beta(x), x \in [b_1, b_2]\}$, when $N$ is large,

$$\dim\{A\} = \frac{(b_2 - b_1)N}{2} + o\left(\frac{(b_2 - b_1)N}{2}\right). \quad (33)$$

The above conclusion can directly apply: when $N$ is large,

$$\dim\{\tilde{A}\} = \frac{ND}{(N - 1)\lambda}(b_2 - b_1) + o\left(\frac{ND}{(N - 1)\lambda}(b_2 - b_1)\right)$$

$$= \frac{DN}{\lambda}(b_2 - b_1) + o\left(\frac{DN}{\lambda}(b_2 - b_1)\right)$$

$$= \frac{M\bar{D}}{\lambda}(b_2 - b_1) + o(M).$$

We can observe that $\dim\{\tilde{A}\}$ has no dependency on $N$. Imagine for any finite aperture $D$, we let $N \to \infty$ so that $D \to 0$. In this case, all elements of $\alpha(x)$ can be seen as $M$ (finite) random samples in the vector $\tilde{\alpha}(x)$. Hence

$$\dim\{B\} \leq \dim\{\tilde{A}\} = \frac{M\bar{D}}{\lambda}(b_2 - b_1) + o(M).$$

Now consider the space $C$. We define $C_q \triangleq \text{span}\{\alpha(x), x \in [b_q^\text{min}, b_q^\text{max}]\}$. An upper bound of its dimension can be obtained by considering the extreme case when all $Q$ spaces are mutually orthogonal so that their dimensions can add up:

$$\dim\{C\} \leq \sum_{q=1}^{Q} \dim\{C_q\} = \sum_{q=1}^{Q} \frac{M\bar{D}}{\lambda}(b_q^\text{max} - b_q^\text{min}) + o(M).$$

Thus, Proposition 2 is proven.

**B. Proof of Theorem 1**

For ease of exposition we omit the user index $k$. Imagine a special case when the scatterers are located in a line which has the length $L$, as shown in Fig. 11. Assume the antennas are far away so that the scatterers are in the same planar wavefront region. We denote the right end of the scattering line as the reference point. The $m$-th antenna is located $\hat{d}_m$ meters away from the reference point, at the angle $\theta_m$. The $p$-th scatter is $\hat{l}(p)$ meters away from the reference point. $\hat{l}(p)$ follows a uniform distribution, i.e., $\hat{l}(p) \sim U(0, \hat{L})$. The phase shift between the scatterer $p$ and the reference point is $2\pi \hat{l}(p) \cos(\theta_m)/\lambda$, $1 \leq m \leq M$. 

![Fig. 9. Uplink sum-rate vs. distance between 2 users, $M = 500$, $r = 15m$, cell-edge SNR 20dB, single-cell network.](image)

![Fig. 10. Uplink per-cell rate vs. $r$, cell-edge SNR 20dB, 7-cell network, each cell has $M = 500$ distributed antennas.](image)
Define a diagonal phase matrix
\[ \Xi \triangleq \text{diag}\{e^{-j2\pi \frac{d_1 + \tau}{\lambda}}, \ldots, e^{-j2\pi \frac{d_N + \tau}{\lambda}}\}. \]

The \( p \)-th scattering path vector channel is now given by:
\[ h_p \triangleq \begin{bmatrix} e^{-j2\pi \frac{l(p) + \tau}{\lambda}} \\ \vdots \\ e^{-j2\pi \frac{|l(p)|\cos(\theta_p)}{\lambda}} \\ e^{j\varphi_p} \end{bmatrix} \Xi \begin{bmatrix} e^{-j2\pi \frac{l(p)\cos(\theta_p)}{\lambda}} \\ \vdots \\ e^{-j2\pi \frac{|l(p)|\cos(\theta_p)}{\lambda}} \\ e^{-j2\pi \frac{l(p)\cos(\theta_p)N}{\lambda}} \end{bmatrix} . \]

If we define \( \tilde{h}_p \triangleq [e^{-j2\pi \frac{l(p)\cos(\theta_p)}{\lambda}}, \ldots, e^{-j2\pi \frac{|l(p)|\cos(\theta_p)}{\lambda}}]^T \), we may see that \( h_p \) and \( \tilde{h}_p \) are unitarily equivalent, since \( e^{j\varphi_p} \Xi \) forms a unitary matrix. If we vary \( \tilde{l}(p) \), we can obtain two linear spaces spanned by \( h_p \) and/or \( \tilde{h}_p \). The dimensions of the two spaces are equal. Therefore in the following we will find the dimension spanned by \( \tilde{h}_p \) instead of \( h_p \). Assume \( N \) is a large integer, we define \( x_n \triangleq -1 + \frac{2(n-1)}{N}, n = 1, \ldots, N \), the set \( \mathcal{X} \triangleq \{x_n\} \), as well as the vector
\[ \mu_p \triangleq \begin{bmatrix} e^{-j2\pi \frac{l(p) + \tau}{\lambda}} \\ e^{-j2\pi \frac{l(p) + \tau}{\lambda}} \\ \vdots \\ e^{-j2\pi \frac{(l(p) + \tau)N}{\lambda}} \end{bmatrix} = e^{j2\pi \tilde{l}(p)} \begin{bmatrix} e^{-j2\pi \frac{l(p)\cos(\theta_p)}{\lambda}} \\ \vdots \\ e^{-j2\pi \frac{l(p)\cos(\theta_p)N}{\lambda}} \end{bmatrix} . \]

Since \( \tilde{l}(p) \sim \mathcal{U}(0, \mathcal{L}) \), we reuse the result shown in (33):
\[ \dim\{\text{span}\{\mu_p, \tilde{l}(p) \in [0, \mathcal{L}]\}\} = 4\frac{\mathcal{L}N}{N\lambda} + o\left(\frac{4\mathcal{L}N}{N\lambda}\right) = \frac{2\mathcal{L}}{\lambda} + o(\mathcal{L}). \]

Note that the above result holds when \( N \) is arbitrarily large. We again observe that when \( N \to \infty \), any \( \cos(\theta_m) \) will fall into the set \( \mathcal{X} \), which indicates
\[ \dim\{\text{span}\{\tilde{h}_p, \tilde{l}(p) \in [0, \mathcal{L}]\}\} \leq \frac{2\mathcal{L}}{\lambda} + o(\mathcal{L}). \]

Recall that the covariance matrix \( R = \mathbb{E}\{\frac{1}{P} \sum_{p=1}^{P} \sum_{q=1}^{P} h_p h_q^H\} \).

Because of the random and independent phases \( \varphi_p \) and \( \varphi_q \),
\[ \forall p \neq q, \mathbb{E}\{h_p h_q^H\} = 0. \]

\[ R = \mathbb{E}\{\frac{1}{P} \sum_{p=1}^{P} h_p h_p^H\} = \mathbb{E}\{h_p h_p^H\}. \]

We can see that the number of scatterers has no impact on the rank of channel covariance matrix. Hence according to (34), the rank of \( R \) is upper bounded by
\[ \text{rank}(R) \leq \frac{2\mathcal{L}}{\lambda} + o(\mathcal{L}). \]

Returning to the one-ring model, we can interpret the ring as the sum of lines, with the total length \( 2\pi r \). An extreme case is when all of the channels corresponding to different pieces of the ring span orthogonal spaces, i.e., the rank of the covariance matrix is the sum of the spatial dimensions corresponding to every pieces of the ring. This is the case when the covariance rank is maximized. Therefore the rank is upper bounded by:
\[ \text{rank}(R) \leq \frac{4\pi r}{\lambda} + o(r). \]

Thus Theorem 1 is proven. \( \square \)

C. Proof of Proposition 2

Proof: We split up the disk centered in the midpoint between the two scatterers into \( N \) (a large number) equal-sized sectors like dividing up a cake. The BS antennas all fall into one of these sectors. We have
\[ ||h_{2q}^H h_{1p}|| = \lim_{N \to \infty} \left| \sum_{n=1}^{N} \sum_{m=1}^{M_n} \left| h_{2qnm} h_{1pnm}^* \right| \right| \]
\[ = \lim_{N \to \infty} \left| \sum_{n=1}^{N} \sum_{m=1}^{M_n} \left( \sqrt{\beta_{2qnm} \beta_{1pnm}} e^{j2\pi \frac{d_{2qnm} - d_{1pnm}}{\lambda}} \right) \right| , \]
where \( M_n \) (could be zero) is the total number of antennas located in the \( n \)-th sector, \( h_{2qnm} \) and \( h_{1pnm} \) are the channel coefficient between the \( m \)-th antenna located in the \( n \)-th sector and the two scatterers. If \( N \) is large, the area contained by the two sides of a sector is small. Therefore BS antennas located in the same sector share the same difference of distances between the two scatterers, i.e., \( d_{2qnm} - d_{1pnm} \) only depends on the sector index \( n \). Based on the fact that \( D_{pq} \) is small, we assume \( \beta_{2qnm} \approx \beta_{1pnm} \) and \( d_{2qnm} - d_{1pnm} = D_{pq} \cos(\alpha_n) \), where \( \alpha_n \triangleq \frac{2\pi r}{N} \) is the AOA from the BS antennas in the \( n \)-th sector.

\[ ||h_{2q}^H h_{1p}|| \approx ||h_{1p}||^2 = \lim_{N \to \infty} \sum_{n=1}^{N} \sum_{m=1}^{M_n} \beta_{1pnm}. \]

Due to the symmetry of the network, when \( M \to \infty \), the radio waves can arrive from any direction with equal probability. Thus \( \sum_{m=1}^{M_n} \beta_{1pnm} \) is independent of sector index \( n \).

\[ \lim_{M \to \infty} \frac{||h_{2q}^H h_{1p}||}{||h_{2q}||} \approx \lim_{N \to \infty} \left| \frac{1}{N} \sum_{n=1}^{N} ||h_{2q}^H h_{1p}|| \right| = \frac{1}{2\pi} \left| \int_{0}^{2\pi} e^{j2\pi \frac{D_{pq} \cos(\alpha)}{\lambda}} d\alpha \right| = J_0\left(\frac{2\pi D_{pq}}{\lambda}\right) , \]
and Proposition 5 is proven.

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