Volume-Outcome Relation for Acute Appendicitis: Evidence from a Nationwide Population-Based Study

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Abstract

Background: Although procedures like appendectomy have been studied extensively, the relative importance of each surgeon’s surgical volume-to-ruptured appendicitis has not been explored. The purpose of this study was to investigate the rate of ruptured appendicitis by surgeon-volume groups as a measure of quality of care for appendicitis by using a nationwide population-based dataset.

Methods: We identified 65,339 first-time hospitalizations with a discharge diagnosis of acute appendicitis (International Classification of Disease, Ninth Revision, Clinical Modification (ICD-9-CM) codes 540, 540.0, 540.1 and 540.9) between January 2007 and December 2009. We used “whether or not a patient had a perforated appendicitis” as the outcome measure. A conditional (fixed-effect) logistic regression model was performed to explore the odds of perforated appendicitis among surgeon case volume groups.

Results: Patients treated by low-volume surgeons had significantly higher morbidity rates than those treated by high-volume (28.1% vs. 26.15, p<0.001) and very-high-volume surgeons (28.1% vs. 21.4%, p<0.001). After adjusting for surgeon practice location, and teaching status of practice hospital, and patient age, gender, and Charlson Comorbidity Index, and hospital acute appendicitis volume, patients treated by low-volume surgeons had significantly higher rates of perforated appendicitis than those treated by medium-volume surgeons (OR = 1.09, p<0.001), high-volume surgeons (OR = 1.16, p<0.001), or very-high-volume surgeons (OR = 1.54, p<0.001).

Conclusion: Our study suggested that surgeon volume is an important factor with regard to the rate of ruptured appendicitis.

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Introduction

Acute appendicitis is the most common surgical emergency. Although the exact mechanism is still obscure, it is likely that luminal obstruction plays a key pathogenic role [1–3]. Luminal obstruction leads to increased mucus production and bacterial overgrowth, which increases the intra-luminal pressure and decreases blood ow to the appendiceal wall. Subsequently, necrosis and perforation of the appendix occur. The primary adverse outcome of appendicitis is appendiceal rupture. In comparison with simple appendicitis, patients with an appendiceal rupture at the time of surgical exploration have been shown to have a higher risk of post-operative complication, such as intra-abdominal abscess, wound infection, and postoperative paralytic ileus (39% vs 8%) [4,5]. It has been assumed that the natural history of appendiceal rupture is within the control of the hospital and attending physician, and that a high rate of rupture reflects a failure of medical care [6]. As a result, appendiceal rupture rates have been proposed as a measure of intrinsic hospital quality, and early surgical exploration has been encouraged to decrease the rupture rate [7,8].

The impact of hospital volume on operative mortality with surgical procedures has been previously studied [9]. One recent review suggested that thousands of preventable surgical morbidities and mortalities occur each year in the United States, and that patients can significantly reduce their risk of operative death by selecting a high-volume hospital [10]. Despite recent interest in surgical volume, many question the applicability of previous research to the volume and outcomes of current practice.

First, due to recent decreases in surgical mortality associated with many procedures [11], it is possible that the relative importance of the volume of procedures performed may be declining. Second, most published studies on volume and outcome have used state-level databases or regional populations that are served by a small number of high-volume centers [12,13]. Although procedures like appendectomy have been studied extensively, the relative importance of each surgeon’s surgical volume to ruptured appendicitis has either not been explored or
has been studied in samples that were too small to permit assessment of performance at meaningful surgical volumes. The purpose of this study was to investigate the rate of ruptured appendicitis with surgeon volume groups as a measure of quality of care for appendicitis by using a nationwide, population-based dataset (from 2007 through 2009). Since some previous studies have suggested that volume-outcome relationships may exist for surgeons but not for hospitals in Taiwan [14,15], the present study focused on the volume-outcome relationship for surgeons. However, we took hospital volume into consideration in the regression model.

### Methods

#### Database

We used data sourced between 2007 and 2009 from the National Health Insurance Research Database (NHIRD). The NHIRD, published by the Taiwan National Health Research Institute, includes all the original claims data and registration files for beneficiaries enrolled under the National Health Insurance (NHI) program. As the end of 2009, about 22.60 of Taiwan’s 22.96 million residents (over 98% coverage rate) were enrolled under the National Health Insurance Research Database (NHIRD). The NHIRD included unique physician identifiers for each medical claim submitted, which enabled us to identify the same physician who performed one or more appendectomies during the three-year study period. Surgeons were sorted in ascending order of appendectomy volume, with volume cutoff points being determined in accordance with prior studies so that the sampled patients were classified into four, approximately equally-sized groups [6,7]. We divided the sample of 63,339 patients into four surgeon volume groups: ≤65 cases (hereafter referred to as low volume), 66–190 cases (medium volume), 121–190 (high volume), and ≥191 cases (very-high volume).

#### Study Sample

We identified 63,339 first-time hospitalizations with a discharge diagnosis of acute appendicitis (International Classification of Disease, Ninth Revision, Clinical Modification [ICD-9-CM] codes 540, 540.0, 540.1 and 540.9) between January 2007 and December 2009. If a patient had ≥2 hospitalizations within a 30-day period, they were regarded as the same episode and we only included the first hospitalization. The mean age for these 63,339 patients was 36.7 (±19.3) years, with a range between 2 and 97 years.

#### Surgeon Acute Appendicitis Volume Groups

The NHIRD included unique physician identifiers for each medical claim submitted, which enabled us to identify the same physician who performed one or more appendectomies during the three-year study period. Surgeons were sorted in ascending order of appendectomy volume, with volume cutoff points being determined in accordance with prior studies so that the sampled patients were classified into four, approximately equally-sized groups [6,7]. We divided the sample of 63,339 patients into four surgeon volume groups: ≤65 cases (hereafter referred to as low volume), 66–190 cases (medium volume), 121–190 (high volume), and ≥191 cases (very-high volume).

#### Key Variables of Interest

The key independent variable of interest was surgeon appendectomy caseload volume. Prior studies have used the proportion of cases of perforated appendicitis as a measure of quality of care for appendicitis [16,17]. Therefore, in this study, we used “whether or not a patient had a perforated appendicitis” as the outcome measure. Patients were regarded as having perforated appendicitis if they had an ICD-9-CM code of 540.0 (acute appendicitis with generalized peritonitis) or 540.1 (acute appendicitis with peritoneal abscess).
Table 2. Distributions of perforated appendicitis by surgeon and patient characteristics and comorbidities (n = 65,339).

| Variable                        | Perforated appendicitis | p value |
|---------------------------------|-------------------------|---------|
|                                 | Yes, n (row %)          | No, n (row %) |
| Surgeon characteristics         |                         |         |
| Surgeon age (years)             |                         |         |
| <41                             | 711 (42.2)              | 20,794 (42.9) | 0.217 |
| 41–50                           | 6,758 (40.1)            | 19,326 (39.9) |
| >50                             | 2,984 (17.7)            | 8,366 (17.2) |
| Surgeon gender                  |                         |         |
| Male                            | 16,284 (96.6)           | 46,848 (96.6) | 0.989 |
| Female                          | 569 (3.4)               | 1,638 (3.4) |
| Practice location                |                         |         |
| Urban                           | 13,323 (79.1)           | 36,874 (76.1) | <0.001 |
| Rural                           | 3,530 (20.9)            | 11,612 (23.9) |
| Teaching status                 |                         |         |
| Yes                             | 15,660 (92.9)           | 43,162 (89.0) | <0.001 |
| No                              | 1,193 (7.1)             | 5,324 (11.0) |
| Patient characteristics         |                         |         |
| Patient age (years)             |                         |         |
| ≤19                             | 3,452 (20.5)            | 10,320 (21.3) | <0.001 |
| 20–39                           | 4,443 (26.4)            | 20,877 (43.1) |
| 40–59                           | 5,024 (29.8)            | 12,477 (25.7) |
| ≥60                             | 3,934 (23.3)            | 4,812 (9.9) |
| Patient gender                  |                         |         |
| Male                            | 9,883 (58.6)            | 25,526 (52.7) | <0.001 |
| Female                          | 6,970 (41.4)            | 22,960 (47.3) |
| Charlson Comorbidity Index score|                         |         |
| 0                               | 14,612 (86.7)           | 45,769 (94.4) | <0.001 |
| 1                               | 1,564 (9.3)             | 2,075 (4.3) |
| 2                               | 307 (1.8)               | 303 (0.6) |
| 3 or more                       | 370 (2.2)               | 339 (0.7) |

Table 1 presents the distribution of the rate of perforated appendicitis and patient and surgeon characteristics across surgeon acute appendicitis case volume groups. Of the 65,339 patients admitted for the treatment of acute appendicitis between January 2007 and December 2009, 16,857 (25.8%) had perforated appendicitis; the rate of perforated appendicitis consistently decreased with increasing surgeon acute appendicitis caseload volume. Patients treated by low-volume surgeons had significantly higher morbidity rates than those treated by high-volume (28.1% vs. 26.15, p < 0.001) or very-high-volume surgeons (28.1% vs. 21.4%, p < 0.001). In addition, patients in the low surgeon volume quartile were more likely to have a greater number of comorbidities (p < 0.001).

Table 1 also shows that patients with acute appendicitis were treated by 2,536 surgeons between 2007 and 2009, at a mean volume per surgeon per 25.8 (±53.2) cases. The x² analyses indicate that the surgeons in the low-volume group were more likely to be female (p < 0.001) and practice in non-teaching hospitals (p < 0.001) than their counterparts in other volume groups. No significant relationships were found in the surgeon distributions in terms of age (p = 0.201) and practice location (p = 0.100) across caseload volume groups.

Table 2 presents the distributions of the rates of perforated appendicitis according to surgeon and patient characteristics. Global x² analyses revealed that there were significant differences in the rate of perforated appendicitis associated with surgeon practice location (p < 0.001), teaching status of practice hospital (p < 0.001), patient age (p < 0.001), patients gender (p < 0.001), and CCI (p < 0.001).

Table 3 provides the crude and covariate-adjusted odds ratio (OR) estimates of the likelihood of perforated appendicitis by surgeon acute appendicitis case volume and surgeon and patient characteristics. The ORs of perforated appendicitis for those patients treated by medium-volume, high-volume, and very-high-volume surgeons were 0.95 (95% CI = 0.90–0.99), 0.91 (95% CI = 0.86–0.95), and 0.70 (95% CI = 0.66–0.73), respectively, compared to patients treated by low-volume surgeons. After adjusting for surgeon practice location, and teaching status of practice hospital, and patient age, gender, CCI, and hospital acute appendicitis case volume, patients treated by low-volume surgeons had significantly higher rates of perforated appendicitis than those treated by medium-volume surgeons (OR = 1.09, reciprocal of 0.92, p < 0.001), high-volume surgeons (OR = 1.16, p < 0.001), or very-high-volume surgeons (OR = 1.54, p < 0.001). In addition, it is also worth noting that higher rates of perforated appendicitis were detected among patients treated by surgeon practicing in the urban areas (adjusted OR = 1.09) and teaching hospitals (OR = 1.61).

Results

Statistical Analysis

We used the SAS package (Version 9.1) for statistical analysis in this study. Chi-squared tests were used to examine differences among surgeon volume groups in terms of both surgeon and patient characteristics. Surgeon characteristics included age, gender, practice location (urban vs. rural), and teaching status of practice hospital, while patient characteristics included age, gender, and Charlson Comorbidity Index (CCI). CCI was used to quantify pre-existing co-morbidities as a means of adjusting for the higher mortality risks associated with co-morbidities (the higher the score, the greater the number of co-morbidities). We then performed a conditional (fixed-effect) logistic regression model (conditional on hospitals in order to partition out systematic hospital-specific variation) to explore the odds of perforated appendicitis among surgeon case volume groups. This conditional model also uses a clustered method for variance estimation to account for the possibility that the patients of each physician have more similar outcomes than patients viewed across physicians. A two-sided p value of less than, or equal to, 0.05 was considered to be statistically significant.

Discussion

To the best of our knowledge, this is the first study to use nationwide data to investigate the relationship between surgeon volume and the perforation rate of appendicitis. Our study demonstrated that low volume surgeons were more likely to have a greater number of perforated appendicitis. Earlier studies have suggested that higher hospital volumes are associated with better outcomes for several surgical procedures [18,19]. The proposed explanation for this inverse volume-outcome relationship is commonly known as the ‘practice makes perfect’ hypothesis. This hypothesis is based upon the rationale that a larger volume of patients provides clinicians with the opportunity to better hone their surgical skills and accumulate...
experience in operation management. Therefore, high-volume providers are more likely to achieve superior clinical performance on account of their greater experience, improved skills, and operational judgment.

In the case of acute appendicitis, the severity and complication rate are closely associated with the time interval between symptom onset and treatment [20], and the rupture rate is generally not the result of surgical technique related error but rather pre- and in-hospital delay. Therefore, high-volume surgeons who have more experience in dealing with patients with acute appendicitis may reduce unnecessary in-hospital delays and lead to better outcomes. A corollary to the “practice makes perfect” hypothesis then, is that low-volume surgeons with poor outcomes would be able to substantially reduce the rate of ruptured appendicitis by simply increasing their patient volumes or having the opportunity to practice under the guidance of more experienced high-volume surgeons.

A second hypothesis used to explain volume-outcome relationships involves ‘selective-referral’. This hypothesis suggests that patients are selectively referred to providers with superior outcomes. Thus, the best providers have the highest volumes on account of their prior record of superior outcomes. However, this hypothesis is unlikely to offer a valid explanation for this study. The progression of appendicitis is rapid, and most patients promptly receive treatment at nearby hospitals. Therefore, ‘selective-referral’ may not be a major factor contributing to the inverse relationship between the patient outcome and surgeon volumes observed in this study.

The main strength of this study was its use of a nation-wide population-based dataset from Taiwan. Taiwan is a country with more than 23 million people on a 35.8 thousand square kilometer island, which launched a universal national health insurance (NHI) program in the end of 1995. Under a single-payer system operated by the Taiwanese government, the NHI program covered 98% of Taiwan’s population by 2009. The population-based dataset associated with this program offered the statistical power necessary to detect real differences. Furthermore, Taiwan is over 98% Han Chinese, and this study could therefore exclude the influence of race and insurance status from our analysis, both of which were previously demonstrated to be important factors influencing the rate of ruptured appendicitis [21,22].

Table 3. Crude and adjusted odds ratios for perforated appendicitis, by surgeon acute appendicitis case volume.

| Variables                          | Crude odds ratio | Adjusted odds ratio |
|------------------------------------|------------------|---------------------|
|                                    | 95% CI           | P value             | 95% CI           | P value             |
| Surgeon acute appendicitis case volume |                  |                     |                  |
| ≤65                                | 1.00             | 1.00                |                  |
| 66–120                             | 0.95 (0.90–0.99) | 0.031               | 0.92 (0.87–0.96) | <0.001              |
| 121–190                            | 0.91 (0.86–0.95) | <0.001              | 0.86 (0.82–0.91) | <0.001              |
| ≥191                               | 0.70 (0.66–0.73) | <0.001              | 0.66 (0.62–0.70) | <0.001              |
| Surgeon characteristics            |                  |                     |                  |
| Practice location                  |                  |                     |                  |
| Urban                              | 1.19 (1.14–1.24) | <0.001              | 1.09 (1.03–1.14) | <0.001              |
| Rural                              | 1.00             | 1.00                |                  |
| Hospital teaching status           |                  |                     |                  |
| Yes                                | 1.62 (1.52–1.73) | <0.001              | 1.61 (1.49–1.73) | <0.001              |
| No                                 | 1.00             | 1.00                |                  |
| Patient characteristics            |                  |                     |                  |
| Patient age (years)                |                  |                     |                  |
| ≤19                                | 1.00             | 1.00                |                  |
| 20–39                              | 0.64 (0.61–0.67) | <0.001              | 0.65 (0.62–0.68) | <0.001              |
| 40–59                              | 1.20 (1.14–1.27) | <0.001              | 1.18 (1.12–1.24) | <0.001              |
| ≥60                                | 2.44 (2.31–2.59) | <0.001              | 2.19 (2.06–2.33) | <0.001              |
| Patient gender                     |                  |                     |                  |
| Male                               | 1.28 (1.23–1.32) | <0.001              | 1.33 (1.28–1.38) | <0.001              |
| Female                             | 1.00             | 1.00                |                  |
| Charlson Comorbidity Index score   |                  |                     |                  |
| 0                                  | 1.00             | 1.00                |                  |
| 1                                  | 2.36 (2.21–2.53) | <0.001              | 1.54 (1.43–1.66) | <0.001              |
| 2                                  | 3.17 (2.71–3.72) | <0.001              | 1.81 (1.53–2.13) | <0.001              |
| 3 or more                          | 3.42 (2.95–3.95) | <0.001              | 2.01 (1.73–2.35) | <0.001              |
| Hospital appendicitis case volume  |                  |                     |                  |
| 1.00 (1.00–1.00)                   | <0.001           | 1.00 (1.00–1.00)    | <0.001           |

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There are some limitations to this study. First, as pathologic confirmation was unavailable in this study, the definition of perforated appendicitis was dependent on ICD-9 codes sourced from an administrative database. These codes may be less accurate than data collected prospectively. For example, while ICD-9-CM...
code 540 is theoretically reserved exclusively for uncomplicated appendicitis, we cannot rule out the possibility that miscoding occurred. Furthermore, as coding is a skill itself, it is possible that lower volume surgeons would have been more likely to miscode more specific codes referring to complicated appendicitis (ICD-9-CM codes 540.0 and 540.1) as ICD-9-CM code 540. If this were the case, our results would be biased toward an underestimation of the effect size. Secondly, although this study has controlled for patient co-morbidities, the administrative database used by this study is extremely limited in its ability to account for the differences in the severity of appendicitis among patients. Therefore it is possible that the associations detected in this study were biased by the residual confounding of severity. Third, this study also lacked any data describing the rate of negative appendectomies which may have biased our results, as higher rates of negative appendectomies would have decreased the rate of perforated appendicitis [23]. Fourth, the dataset used in this study is extremely limited in its ability to account for differences in the severity of appendicitis among patients. Therefore it is possible that the associations detected in this study were biased by the residual confounding of severity. Third, this study also lacked any data describing the rate of negative appendectomies which may have biased our results, as higher rates of negative appendectomies would have decreased the rate of perforated appendicitis [23].

In conclusion, our study suggested that surgeon volume is an important factor with regard to the rate of ruptured appendicitis. Future investigations will need to be undertaken to determine a threshold of surgeon volume for best practice, and to identify the differences in clinical approach between high-volume surgeons with superior outcomes, and low-volume surgeons with inferior outcomes. The identification of such differences may be useful toward aiding low-volume surgeons with inferior outcomes to reduce the incidence of perforated appendicitis.

**Author Contributions**

Conceived and designed the experiments: HCL, PLW. Analyzed the data: HCL, SPL. Contributed reagents/materials/analysis tools: HCL. Wrote the paper: HCL, PLW, SPL, JJK. Critical revision: HCL, PLW, SPL, JJK.

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