Target temperature management following cardiac arrest: a systematic review and Bayesian meta-analysis

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Abstract
Background: Temperature control with target temperature management (TTM) after cardiac arrest has been endorsed by expert societies and adopted in international clinical practice guidelines but recent evidence challenges the use of hypothermic TTM.

Methods: Systematic review and Bayesian meta-analysis of clinical trials on adult survivors from cardiac arrest undergoing TTM for at least 12 h comparing TTM versus no TTM or with a separation > 2 °C between intervention and control groups using the PubMed/MEDLINE, EMBASE, CENTRAL databases from inception to 1 September 2021 (PROSPERO CRD42021248140). All randomised and quasi-randomised controlled trials were considered. The risk ratio and 95% confidence interval for death (primary outcome) and unfavourable neurological recovery (secondary outcome) were captured using the original study definitions censored up to 180 days after cardiac arrest. Bias was assessed using the updated Cochrane risk-of-bias for randomised trials tool and certainty of evidence assessed using the Grading of Recommendation Assessment, Development and Evaluation methodology. A hierarchical robust Bayesian model-averaged meta-analysis was performed using both minimally informative and data-driven priors and reported by mean risk ratio (RR) and its 95% credible interval (95% CrI).

Results: In seven studies (three low bias, three intermediate bias, one high bias, very low to low certainty) recruiting 3792 patients the RR by TTM 32–34 °C was 0.95 [95% CrI 0.78—1.09] for death and RR 0.93 [95% CrI 0.84—1.02] for unfavourable neurological outcome. The posterior probability for no benefit (RR ≥ 1) by TTM 32–34 °C was 24% for death and 12% for unfavourable neurological outcome. The posterior probabilities for favourable treatment effects of TTM 32–34 °C were the highest for an absolute risk reduction of 2–4% for death (28–53% chance) and unfavourable neurological outcome (63–78% chance). Excluding four studies without active avoidance of fever in the control arm reduced the probability to achieve an absolute risk reduction > 2% for death or unfavourable neurological outcome to ≤ 50%.

Conclusions: The posterior probability distributions did not support the use of TTM at 32–34 °C compared to 36 °C also including active control of fever to reduce the risk of death and unfavourable neurological outcome at 90–180 days. Any likely benefit of hypothermic TTM is smaller than targeted in RCTs to date.

Keywords: Cardiac arrest, Target temperature management, Bayesian statistics

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practice guidelines [1–3] based on several randomised controlled trials (RCTs), systematic reviews and meta-analyses [4–10]. The original studies and systematic reviews have exclusively used frequentist methods for statistical inference. A growing body of literature is instead using Bayesian statistics to interpret the results of clinical trials, i.e. the likelihood, in the context of already existing beliefs on treatment effects, i.e. the prior, that when combined generate a posterior distribution of probabilities for the effect size [11–13]. This brings the trial effect estimates beyond the dichotomous outcome as being significant or non-significant based on a p value of 0.05 in frequentist hypothesis testing and instead attributes probabilities to effect sizes contained in the 95% credible interval. Bayesian analyses have been proposed to supplement interpretations of clinical trials in critical care [14], to inform clinical practice guidelines in cardiology [15] and the US Food and Drug Administration has issued guidelines for their use in clinical trials of medical devices [16].

The recently published trial by Dankiewicz et al. [17], the largest study performed to date on the use of TTM in post-cardiac arrest care, compared the institution of TTM at 33–34 °C for 24 h with active measures to maintain normothermia (<37.5 °C) if the body temperature increased above 37.8 °C. The study was powered to detect an absolute risk reduction for death by 7.5% by frequency inference and did not observe any benefit of hypothermic TTM on all-cause mortality or poor functional recovery. The aim of this review and Bayesian meta-analysis was to evaluate the effect of TTM on survival and neurological outcome compared to no TTM or avoiding pyrexia in the care of adult, comatose survivors of cardiac arrest. The null hypothesis of no difference between patients treated with TTM at 32–34 °C or TTM ≥ 36 °C was compared to the alternative hypothesis that hypothermic TTM confers a benefit using a range of data-informed priors to reflect optimism, pessimism or equipoise regarding the effect of TTM. By using Bayesian statistical inference, the posterior probabilities attributed to the absence of any benefit could be compared to a range of treatment effects, including ones smaller than targeted in the original studies.

Methods

This review and meta-analysis were performed in accordance with the protocol registered with PROSPERO (CRD42021248140) and are reported according to the PRISMA statement [20] (Additional File 1: Table S1) and the ROBUST criteria [21] (Additional File 1: Table S2).

Eligibility criteria

All randomised and quasi-randomised (e.g. allocation based on the day of week) controlled trials of adult (≥ 18 years of age) comatose survivors from cardiac arrest undergoing TTM for at least 12 h were included. Hence, both in-hospital (IHCA) and out-of-hospital cardiac arrest (OHCA) with all initial rhythms and any locations to initiate TTM were considered. A separation between intervention and control groups in TTM studies was accepted as any TTM temperature compared to no TTM or a difference >2 °C in target temperature between groups. The intervention and control groups are referred to by the target temperature, i.e. TTM32–34 versus TTM≥36. We excluded studies comparing similar TTM but of different duration, investigating hypothermic (32–34 °C) TTM in both interventional and control arms, using different devices to induce TTM or studies focusing on the setting of TTM, e.g. pre-hospital versus in-hospital studies.

Information sources and search strategy

The PubMed/MEDLINE, EMBASE, CENTRAL bibliographic databases and the clinicaltrials.gov trial database were searched using a search strategy directed towards randomised controlled trials in human adults from inception to the 1 September 2021 (original search, Additional File 1: Table S3) with an expanded search performed during editorial review (last updated 24 January 2022, Additional File 1: Table S4). Search results were uploaded to Covidence (Veritas Health Innovation, Melbourne, Australia). Only studies providing the essential data were considered. Titles and abstracts were screened for potential eligibility with the full text reviewed in case of unclear potential. Citations from included articles were also reviewed as were citations in published systematic reviews and meta-analyses.

Data items collected, risk of bias and certainty of evidence

A standardised data abstraction form was used to capture author, publication year, cardiac arrest characteristics, details of temperature management in control and intervention groups, the number of events and patients in in each group and times when outcomes were censored. Survival (primary outcome) and neurological status (secondary outcome) using the definitions in the original studies were censored within up to 180 days after cardiac arrest. Most neurocognitive recovery occurs within the first 90–180 days timeframe [22, 23] that has commonly been used in RCTs of TTM. Short-term neurocognitive outcome, e.g. at discharge from intensive care or from hospital, might be susceptible to confounding by the time of discharge reflecting the speed and quality of recovery following cardiac arrest. These timeframes are
reported separately in this review. Domains of bias were assessed using the updated Cochrane risk-of-bias (RoB 2) for randomised trials tool [24] and certainty of evidence assessed using the Grading of Recommendation Assessment, Development and Evaluation (GRADE) methodology [25]. The search, study selection, data extraction, bias and GRADE assessments were performed independently by two authors (AA and MS), and any discrepancies were resolved by consensus involving a third researcher (SF).

**Effect measures, data synthesis and sensitivity analyses**

The risk ratio (RR) and its upper and lower 95% confidence interval (95% CI) for death and unfavourable neurological outcome were retrieved or calculated for each study. Bayesian meta-analysis was used to examine the studies in aggregate and sequentially, i.e. using incremental evidence to generate posterior probabilities for the fixed-effect and random-effect models [26] of the null hypothesis ($H_0$, there is no difference between TTM$_{32–34}$ and TTM$_{≥36}$), illustrated by the Bayes factor BF$_{01}$, a likelihood ratio in favour of $H_0$) and the alternative hypothesis ($H_1$, TTM$_{32–34}$ results in a difference compared to TTM$_{≥36}$, illustrated by the Bayes factor BF$_{10}$, a likelihood ratio in favour of $H_1$). The likelihood ratio contained in the Bayes factor (BF) represents a metric for the strength of supporting evidence. A hierarchical robust Bayesian model-averaged meta-analysis [27] was performed that combined the results of Bayesian fixed-effect and Bayesian random-effect models to generate a model average for effect size, heterogeneity and publication bias [28]. A range of priors were set for the effect size with the mean representing the belief where the treatment effect is centred (RR $< 1$, TTM$_{32–34}$ confers benefit; RR $= 1$, no effect; RR $≥ 1$ TTM$_{32–34}$ does not confer benefit) and the variance (standard deviation) representing the certainty in the belief (a certain belief has a narrower variance compared to an uncertain belief) (Additional File 1: Table S4). A minimally informative effect size prior was set as RR $= 1$ (no effect of intervention) with a Cauchy distribution of probabilities (i.e. a continuous distribution of probabilities) meaning that all information is provided by the studies. The minimally informed prior hence produces results most similar to a frequentist meta-analysis. In addition, effect size priors informed by meta-analysis or reflecting a range from strongly pessimistic beliefs about plausible TTM outcomes were used [29]. Frequentist meta-analysis of binary outcomes by random effects estimates using the DerSimonian–Laird method was performed to generate informed priors (Additional File 1: Fig. S1). Data-driven priors for between-study heterogeneity [30] were used together with priors for publication bias. All prior settings are reported in Table 1). The model-average overall effect is reported as the mean RR including its 95% credible interval (95% CrI), i.e. the interval within which there is a 95% probability the treatment effect resides. The Markov chain Monte Carlo (MCMC) algorithm (3 chains, 10,000 iterations, burn-in of 5000 iterations, adaptation of 1000 iterations, thinning of 1, target margin of error 1%, target Rhat $< 1.05$, exclude models with Rhat $> 1.05$) was used to derive posterior effect estimates and 95% CrIs. Convergence of the MCMC chains was assessed by Rhat at 1, visual inspection of the trace and density plots to confirm that the chains were producing representative values from the posterior distribution. Finally, the autocorrelation plots were inspected to demonstrate essentially zero autocorrelation. Various treatment effects of TTM were evaluated based on the posterior probabilities for the overall effect size. First, considering the frequentist approach to refute the null hypothesis ($H_0$), any decrease in mortality (RR $< 1$) was considered. Second, an absolute risk reduction (ARR) in mortality by 7.5% was explored as per the most recent trial protocol [18, 19]. Third, the incidence of OHCA in the US [31] and Europe [32] provide a crude theoretical estimate of 300,000 OHCA patients annually meaning that a significantly smaller ARR of 2% would still translate into 6000 lives saved every year and arguably represents a minimum clinically important difference [33]. Fourth, a range of ARR at 4%, 6% and 10% was included to illustrate treatment effects below and above the protocol effect size estimate. The baseline was set using data from the International Cardiac Arrest Registry and a separation of TTM aligned with this review [34]. Fifth, the posterior probability for no benefit of TTM was assessed by a RR $≥ 1$. Sensitivity analyses were performed for studies that only reported short-term neurological outcome, did or did not apply an explicit temperature definition of normothermia and the avoidance of fever in the control groups and for studies that reported patients with initial shockable versus non-shockable rhythms separately.

All analyses were performed using RStudio (version 1.3.1093) [35] and JASP (version 0.14.1) [36] including the meta, RoBMA, metaBMA, MCMCpack packages.

**Results**

The original and expanded search strategies generated 1040 and 4201 unique publications, respectively, of which 186 and 80 were assessed for full-text eligibility with the same 7 trials that recruited 3792 patients included in the final analyses (Fig. 1; Additional File 1: Table S4 bottom) [17, 37–42] of which 1898 received TTM. One study [42] reported neurological function at hospital discharge. A temperature of 32–34 °C was used in the intervention groups (TTM$_{32–34}$). Patients in the control groups lacked explicit temperature definitions of normothermia.
in four studies [37–39, 42] while TTM at 36–37.5 °C and avoiding fever was applied in three studies [17, 40, 41] (TTM ≥ 36). The majority of patients were OHCA with IHCA only included in two studies [37, 41]. The characteristics of included studies are reported in Table 2.

Three studies were assessed as having low risk of bias, three studies with some concerns relating to the unblinding of the intervention and lack of pre-published protocols or statistical analysis plans and one with further high risk given its quasi-randomised design (Fig. 2). The certainty of evidence was graded as very low for survival and low for neurological outcome (Additional File 1: Table S5).

The posterior likelihood favoured the null-hypothesis (H₀) of no difference in deaths between TTM32–34 and TTM ≥36 groups with BF₀₁ = 26.1, i.e. the evidence for H₀ is 26 times more likely than for H₁ (Fig. 3A). The final posterior probability of the null hypothesis with sequential addition of studies was 82% in a fixed effects model and 14% in a random effects model (Fig. 3C). Conversely, the posterior probability supporting the hypothesis of benefit by TTM32–34 was 3% in a fixed and 1% in a random effects model (Fig. 3C). The posterior likelihood favoured the null hypothesis of no difference in neurological outcome between TTM32–34 and TTM ≥36 groups with BF₀₁ = 6.67, i.e. the evidence for H₀ is 7 times more likely than for H₁ (Fig. 3B). Incremental study data generated a posterior probability for the null hypothesis of 53% and 28% in fixed and random models, respectively, with a posterior probability for the hypothesis of benefit by TTM32–34 at 7% in a fixed and 2% in a random model (Fig. 3D).

The Bayesian meta-analysis of survival demonstrated a mean RR for death 0.96 (95% CrI 0.82–1.04) (Fig. 4A), with evidence for heterogeneity (r), BF₁ = 225, but without evidence of publication bias, BF₁ = 1.24. The Bayesian meta-analysis of neurological recovery demonstrated a mean RR for unfavourable neurological outcome 0.93

### Table 1

Settings of the data-driven effect size, study heterogeneity and publication bias priors for the primary outcome mortality and secondary outcome unfavourable neurology

| Prior description                                      | Risk ratio, mean | Risk ratio, standard deviation |
|--------------------------------------------------------|------------------|-------------------------------|
| **Death**                                              |                  |                               |
| Minimally informative                                   | Cauchy distribution with location = 1 | NA                           |
| Informed based on frequentist random effects meta-analysis | 0.95             | 0.04                          |
| Strongly enthusiastic based on an RR as targeted in the TTM2 study [17] with a SD similar to the TTM study [40] | 0.86             | 0.07                          |
| Moderately enthusiastic based on an RR similar to Lascarrou et al. [41] with a SD similar to the TTM study [40] | 0.98             | 0.07                          |
| Moderately sceptic based on an RR= 1 with a SD similar to the TTM study [40] | 1                | 0.07                          |
| Strongly sceptic based on an RR= 1 with a SD half that of the TTM study [40] | 1                | 0.035                         |
| **Unfavourable neurological outcome**                  |                  |                               |
| Minimally informative                                   | Cauchy distribution with location = 1 | NA                           |
| Informed based on frequentist random effects meta-analysis | 0.94             | 0.04                          |
| Strongly enthusiastic based on an RR as targeted in the TTM2 study [17] with a SD similar to the TTM study [40] | 0.86             | 0.06                          |
| Moderately enthusiastic based on an RR similar to Lascarrou et al. [41] with a SD similar to the TTM study [40] | 0.95             | 0.06                          |
| Moderately sceptic based on an RR= 1 with a SD similar to the TTM study [40] | 1                | 0.06                          |
| Strongly sceptic based on an RR= 1 with a SD half that of the TTM study [40] | 1                | 0.03                          |

Based on 14,886 meta-analyses [26]:

| Between-study heterogeneity (r²)                      | Mean | Standard deviation |
|-------------------------------------------------------|------|--------------------|
| Outcome: death                                        | 0.02 | 0.51               |
| Outcome: unfavourable neurological recovery           | 0.02 | 1.23               |

### Publication bias

| Two-sided p value < 0.05                              | 100% publication rate |
| Two-sided p value ≥ 0.05 < 0.10                       | 66% publication rate  |
| Two-sided p value ≥ 0.10                              | 33% publication rate  |

*Selection models use weighted distributions to account for the proportion of studies that are missing because they yielded non-significant results*
(95% CrI 0.84–1.02) (Fig. 4B), again with evidence for heterogeneity ($\tau$), BF = 308 but not for publication bias, BF = 1.11. These RRs remained unchanged in the sensitivity analyses using a range of effect size priors (Additional File 1: Table S6). In the sensitivity analyses excluding studies without explicit temperature definition of normothermia and avoidance of fever in the control group, the heterogeneity was substantially reduced (BF = 33 for survival and BF = 35 for neurological outcome). The RR for death changed to 0.99 (95% CrI 0.69–1.14) (Fig. 4C) and the RR for unfavourable neurological outcome to 0.96 (95% CrI 0.68–1.12) (Fig. 4D).

The posterior probabilities for treatment effects at different thresholds are given in Table 3 for death and in Table 4 for neurological outcome.

The posterior probability of no benefit conferred (RR $\geq$ 1) by TTM$_{32-34}$ compared to TTM$_{\geq 36}$ was 24% and 12% for death and unfavourable neurological outcome, respectively. Minor treatment effects of TTM$_{32-34}$ at ARR of 2–4% had the greatest posterior probabilities (28–78%). Including short-term neurological outcome at hospital discharge increased the posterior probabilities of benefit by TTM$_{32-34}$ and extended the range where chances of benefit outweighed no-benefit up to
| Study       | Study setting          | Years     | Inclusion criteria                                                                 | Number of patients | Intervention group                                                                 | Control group                                                                 | Primary outcome                                                                 |
|-------------|------------------------|-----------|-------------------------------------------------------------------------------------|--------------------|-------------------------------------------------------------------------------------|---------------------------------------------------------------------------------|--------------------------------------------------------------------------------|
| Bernard     | Single centre in Melbourne, Australia | 1996–1999 | OHCA, initial rhythm VF, age > 18 (men) and > 50 (women)                             | 43                 | TIM at 33 °C for 18 h                                                                | No TTM, unclear about fever treatment                                           | Proportion of CPC 1–2 at hospital discharge                                     |
| HACA        | Multiple centres in Europe | 1996–2001 | OHCA and some IHCA, shockable rhythm, witnessed, age 18–75                          | 275                | TIM at 32–34 °C, for 24 h                                                            | No TTM, unclear about fever treatment                                           | Proportion of CPC 1–2 at six months                                              |
| Laurent     | Single centre in Paris, France | 2000–2002 | OHCA only, all rhythms, age 18–75                                                   | 42                 | TIM at 32–33 °C for 16 h                                                             | No TTM, unclear about fever treatment                                           | Survival until six months                                                      |
| Hachimi-Idrissi | Single centre in Belgium  | 1999–2002 | OHCA only, non-shockable rhythm, age > 18 years                                      | 28                 | TIM at 33 °C for 24 h                                                                | No TTM, unclear about fever treatment                                           | Level of s100b                                                                 |
| Nielsen     | Multiple centres in Europe and Australia | 2010–2013 | OHCA, all rhythms, age > 18                                                         | 950                | TIM at 33 °C for 28h                                                                 | TIM at 36 °C for 28h                                                            | All-cause mortality until six months                                            |
| Lascarrou   | Multiple centres in France         | 2014–2018 | OHCA or IHCA, non-shockable rhythm, age > 18                                         | 584                | TIM at 33 °C for 24 h                                                                | TIM at 36.5–37.5 °C                                                            | Proportion of CPC 1–2 at 90 days                                                |
| Dankewicz   | Multiple centres in Europe, Australia, New Zealand and USA | 2017–2020 | OHCA, all rhythms, age > 18                                                         | 1900               | TIM at 33 °C for 24 h                                                                | TIM, if > 37.8 °C then 37.5 °C                                                | All-cause mortality until six months                                            |

*OHCA* out-of-hospital cardiac arrest, *VF* ventricular fibrillation, *TTM* targeted temperature management, *HACA* hypothermia after cardiac arrest, *IHCA* in-hospital cardiac arrest, *CPC* cerebral performance category
ARR > 10%. The probability for no benefit of TTM$_{32-34}$ (RR $\geq$ 1) increased to 45% for death and to 37% for unfavourable neurological outcome with the probabilities for achieving an ARR $\leq$ 50% and for ARR $\leq$ 1% less than the chance of no benefit when only studies with explicit temperature definition of normothermia and avoidance of fever in the control groups were considered. In patients with an initial shockable versus non-shockable rhythm, the RR for death was 0.91 (95% CrI 0.66–1.06) versus 1.00 (95% CrI 0.93–1.07) (Additional File 1: Fig. S2) and for unfavourable neurological outcome 0.87 (95% CrI 0.60–1.04) versus 0.89 (95% CrI 0.57–1.09) (Additional File 1: Fig. S3).

Discussion
This systematic review of randomised controlled trials of TTM in post-cardiac arrest care used Bayesian meta-analysis to assess posterior probabilities for a range of effect sizes by hypothermia (TTM$_{32-34}$) compared with normothermia (TTM$_{\geq 36}$) with or without avoidance of fever on survival and neurological recovery. The likelihood for the hypothesis of no difference between TTM$_{32-34}$ and TTM$_{\geq 36}$ was 26 times that of benefit by TTM$_{32-34}$ for death and 8 times for unfavourable neurological recovery. The posterior probability of a 7.5–10% absolute risk reduction by TTM$_{32-34}$ as targeted in the included trials was < 20% and only the posterior probabilities for achieving a 2% absolute risk reduction for unfavourable neurological outcome.

A Bayesian meta-analysis was chosen for this review as it can incorporate external data such as results from other meta-analyses or clinicians’ views on effect sizes. It can also demonstrate the relationship between treatment

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**Risk of bias assessment. [23]**

| Study                  | Bias arising from the randomisation process | Bias due to deviation from intended interventions | Bias due to missing outcome data | Bias in measurement of outcome | Bias in selection of reported results | Overall |
|------------------------|-----------------------------------------------|-----------------------------------------------|---------------------------------|---------------------------------|--------------------------------------|---------|
| Bernard 2002 [41]     | High\textsuperscript{a}                      | Some concerns\textsuperscript{b}               | Low                             | Low                             | Some concerns\textsuperscript{c}     | High    |
| HACA 2002 [36]        | Low                                           | Some concerns\textsuperscript{b}               | Low                             | Low                             | Some concerns\textsuperscript{c}     | Some concerns |
| Laurent 2005 [37]     | Low                                           | Some concerns\textsuperscript{b}               | Low                             | Low                             | Some concerns\textsuperscript{c}     | Some concerns |
| Hachimi-Idrissi 2005 [38] | Low                                      | Some concerns\textsuperscript{b}               | Low                             | Low                             | Some concerns\textsuperscript{c}     | Some concerns |
| Nielsen 2013 [39]     | Low                                           | Some concerns\textsuperscript{b}               | Low                             | Low                             | Low                                  | Low     |
| Lascarrou 2019 [40]   | Low                                           | Some concerns\textsuperscript{b}               | Low                             | Low                             | Low                                  | Low     |
| Dankiewicz 2021 [17]  | Low                                           | Some concerns\textsuperscript{b}               | Low                             | Low                             | Low                                  | Low     |

*Fig. 2* Risk of bias assessment for the included studies using the updated Cochrane RoB2 tool for randomised trials [23]. For each of the bias domains, the risk was classified as ‘low risk,’ ‘some concerns’ or ‘high risk.’ \textsuperscript{a}Randomisation by odd or even date of month; \textsuperscript{b}Intervention not blinded to treating clinicians; \textsuperscript{c}No published pre-specified statistical analysis plan.
risk and benefit [43] within a limited number of studies [44] as the observed treatment effect of one trial is informed by that of the other trials. The aggregate of the estimated intervention effect is thus less susceptible to trials with small or extreme results. The 95% credible interval of posterior probabilities allows for a range of treatment effects to be evaluated which might be attractive to clinicians considering individualised TTM in post-cardiac arrest patients influenced by idiosyncratic factors associated with presence or absence of any benefit.

The rejection or acceptance of a hypothesis model is common to scientific inquiry. This study evaluated the null hypothesis (H0) of no difference between patients treated with TTM at 32–34 °C compared to TTM ≥ 36 °C with or without treating fever versus the alternative hypothesis (H1) that hypothermic TTM confers a benefit. The Bayesian analysis illustrated the likelihood of the body of evidence considered in the rejection or acceptance of the TTM hypothesis, notwithstanding the very low to low certainty in the evidence. The incremental study data provided strong evidence (BF = 26) in support of H0 for death and substantial evidence (BF = 8) for H0 regarding neurological outcome. The posterior probabilities for the alternative hypothesis (H1) of benefit by TTM32–34 were < 10% in both the fixed effect, i.e. all studies share a common treatment effect, and the random effect, i.e. studies share a distribution of treatment effects, models. There is thus an absence of evidence to support TTM32–34 for survival given the 1–3% posterior probability. Furthermore, there is plausible absence of evidence for neurological outcome based on the 2–7% posterior probability. It must be noted that the posterior probability in the random effects model for the hypothesis of equipoise (H0) was around 25% for both outcomes in the random effects model that is less convincing and suggests further trials are needed comparing TTM32–34 with other temperatures.

Previous frequentist meta-analyses of TTM have reported an overall (considering all studies regardless of initial rhythm, TTM characteristics and times for outcome assessment) non-significant effect on survival and a variable but mostly favourable effect on neurological outcomes by TTM32–34 [6, 7, 9, 10, 45]. Reviews of TTM in post-cardiac arrest care using expanded inclusion
criteria and including retrospective, observational cohort studies, while unsuitable to inform clinical practice, have reported a benefit of TTM. Two recent meta-analyses including the TTM2 trial concluded that various levels of hypothermic TTM may not improve survival or neurological outcome compared to normothermia while associated with higher incidence of arrhythmias. The Bayesian meta-analysis used in this review allowed for extended observations. First, the posterior probability distribution for death and unfavourable neurological outcome beyond a risk ratio of 1 indicated when TTM may be at best futile and at worst could potentially be associated with harm. These probabilities (24% for death and 12% for unfavourable neurological outcome in all studies and 45% and 37%, respectively, in the sensitivity analysis) warrant careful consideration. The model-averaged effect (overall) in bold black text and black diamond. RR risk ratio, 95% CrI 95% credible interval for the posterior probability distribution.

### Table 3

Posterior probability (%) of treatment effect by specified threshold criteria for death for all studies (top) and only for studies with explicit temperature definition of normothermia and using fever avoidance in the control group (bottom)

| No benefit RR ≥ 1 | Any benefit RR < 1 | ARR > 2% | ARR > 4% | ARR > 6% | ARR 7.5% | ARR > 10% |
|--------------------|---------------------|--------|--------|--------|--------|--------|
| Death up to 180 days [17, 36–41] | | | | | | |
| 24 | 76 | 53 | 28 | 11 | 4 | 1 |
| Death up to 180 days Explicit temperature definition of normothermia and using fever avoidance in control group [17, 39, 40] | | | | | | |
| 45 | 55 | 42 | 30 | 20 | 14 | 7 |

RR risk ratio, ARR absolute risk reduction.
is not devoid of adverse effects, e.g. arrhythmias, prolonged mechanical ventilation and extended stay in ICU [17, 34]. Second, the posterior probabilities of benefit by RR < 1 were highest for an ARR of 2–4%. There was a 28–67% chance of achieving this result by TTM32–34 for death and unfavourable neurological outcome in all studies but this was reduced to 30–50% in the sensitivity analysis. Such magnitudes of treatment effects might not be discernible or convincing in the local context of providing care to comatose survivors of cardiac arrest but could still carry benefits on a hospital network or population level. An international survey of individuals involved in cardiac resuscitation reported a minimum clinically important difference of 2–6% for survival with good neurological outcome at hospital discharge across different cardiac arrest characteristics [33]. However, this Bayesian analysis demonstrated that at a level where 25 patients need to be treated to gain one survivor, the chance of benefit by TTM32–34 is almost on par with the absence of any benefit. This challenges the routine use of hypothermic TTM. In terms of neurological outcome at a similar risk reduction level the balance of probabilities was in favour of TTM32–34 but arguably still not forming a compelling case for routine hypothermic TTM. Third, treatment effects by TTM32–34 at an ARR of 7.5–10%, while typically targeted in the sample size calculations for the included RCTs, had very low chances (1–17%) of being achieved. Fourth, in sensitivity analyses for a range of effect size priors including strongly enthusiastic as well as sceptic beliefs on TTM, the RR and 95% CrI remained principally unchanged and support the overall RR estimate as robust. The posterior probabilities of benefit by TTM32–34 on survival and neurological recovery were however shifted towards benefit in patients with an initial shockable compared to non-shockable rhythm although this difference was substantially reduced if studies not using a specific temperature to define normothermia in the control group were removed.

Study heterogeneity in previous [4–10] and recent [47–50] systematic reviews and meta-analyses has been moderate to high [51] and largely related to studies where hypothermia was compared to no temperature control including a proportion of patients with febrile temperatures. The addition of the recent study by Dankiewicz et al. [17] to the cumulative evidence has reduced heterogeneity given its sample size and protocolised avoidance of fever, similar to the original study by Nielsen et al. [40]. Notably, approximately 40% of patients in the study by Dankiewicz et al. [17] required active temperature management to maintain normothermia. The Bayesian meta-analysis demonstrated an approximate ten-fold reduction in the posterior probability of heterogeneity by excluding studies without explicit temperature definition of normothermia and avoiding fever and justified the sensitivity analysis of the remaining studies. The evidence still supported remaining heterogeneity and differences in the design, conduct and reporting of studies as well as general post-cardiac arrest care are essential considerations in the translation of research results into practice. It seems that clinicians may have interpreted the frequentist dichotomy driven by a threshold p value as evidence of absence of a treatment effect in non-significant (p > 0.05) TTM studies. The Nielsen et al. trial that reported no significant difference between TTM at 33 and 36 °C [40] was followed by a decreased use of TTM with increased incidence of febrile temperatures in cardiac arrest patients admitted to ICU [52–54]. It would seem prudent to avoid a similar dismissal of TTM [55] after the most recent trial not at least considering that the posterior probabilities for any benefit by hypothermic TTM still remained greater than the risk of no benefit in the sensitivity analyses. The importance of fever prevention also warrants further investigation. Several aspects of TTM remain the subject of debate and ongoing clinical research, including patient selection (e.g. shockable vs. non-shockable rhythms), timing and speed of instituting

| Table 4 | Posterior probability (%) of treatment effect by specified threshold criteria for unfavourable neurological outcome censored at 90–180 days following cardiac arrest (top) or between hospital discharge to 180 days (middle) |
|---|---|---|---|---|---|---|
| Unfavourable neurological outcome at 90–180 days [17, 36–40] | Any benefit RR < 1 | ARR > 2% | ARR > 4% | ARR > 6% | ARR 7.5% | ARR > 10% |
| 17 | 83 | 67 | 47 | 28 | 17 | 6 |
| Unfavourable neurological outcome, hospital discharge to 90–180 days [17, 36–41] | 12 | 88 | 78 | 63 | 45 | 32 | 14 |
| Unfavourable neurological outcome, 90–180 days explicit temperature definition of normothermia and using fever avoidance in control group [17, 39, 40] | 37 | 63 | 50 | 36 | 25 | 17 | 9 |

The bottom row only includes studies with explicit temperature definition of normothermia and using fever avoidance in the control group (bottom) RR risk ratio, ARR absolute risk reduction
TTM, the optimal target temperature, duration, mode of reducing body temperature and management of the rewarming phase. The RCTs to date have been underpowered to detect the most likely effect size estimates of 2–4% demonstrated in this meta-analysis. A future RCT aimed at this effect using frequentist statistical inference would require prohibitively large study populations with between 2500 and 9000 participants per arm of the trial. Adaptive platform trial designs that employ Bayesian statistical inference might prove more realistic. Such future adaptive trials could investigate multiple domains of TTM management and assign more enthusiastic priors to early initiation of TTM, shockable rhythms, neurological outcomes with a sceptic prior for hypothermic TTM and an informed prior for TTM ≥36 compared to higher temperatures. Enrichment strategies could be used to identify patients for whom a reduced risk by 2–6% is meaningful.

**Strengths and limitations**
This systematic review and meta-analysis is strengthened by incorporating the study by Dankiewicz et al. [17] that doubles the study population compared to the previous literature. The use of Bayesian statistical inference to evaluate TTM has not been reported before and allowed treatment effects to be evaluated on a continuous spectrum of probabilities. The informed prior was aligned with the results of recent frequentist systematic reviews and meta-analyses [47, 48] and may thus be considered to reflect best current evidence. Several important limitations should be noted. The analysis comprised studies conducted predominantly in European hospitals. Other synonyms/ permutations in the search strings might be considered although no additional studies relevant to this analysis have been identified in other recent systematic reviews using other comprehensive search strategies [47–50]. The study by Dankiewicz et al. [17] contributed 50% of the patient cohort and is hence influential on the results. Other important patient-centric outcomes such as quality of life or cognitive function and outcomes beyond 180 days were not included. This review compared hypothermic TTM at 32–34 °C with TTM at ≥36 °C and thus variable effects within the hypothermic range, e.g. 31.5 versus 34 °C (CAPITAL-CHILL, NCT02011568) or 32 versus 34 °C [56] were not explored. The results are influenced by the variable quality of included studies and neither adjunct therapies, details of TTM management nor determination of futility could be consistently assessed. While data-driven priors were used in this study, other settings for priors may be used and could generate different posterior distributions of probabilities [43, 57, 58]. In settings where the baseline incidence of death or unfavourable neurological outcome is higher or lower, the thresholds in the distribution of posterior probabilities for the ARR illustrations would correspond to higher and lower chances, respectively, to achieve similar reductions in RR.

**Conclusions**
This Bayesian meta-analysis of randomised controlled trials of target temperature management (TTM) for at least twelve hours in adult comatose survivors of cardiac arrest did not support the use of TTM at 32–34 °C as compared to ≥36 °C also including active control of fever, to reduce the risk of death and unfavourable neurological outcome at 90–180 days. Future studies would need to consider that the most probable effect size estimates of TTM are less than half of that targeted in trials to date.

**Abbreviations**
95% CI: 95% Confidence interval; 95% CrI: 95% Credible interval; ARR: Absolute risk reduction; GRADE: Grading of Recommendation Assessment, Development and Evaluation; RCT: Randomised controlled trial; RoB: Risk of bias; RR: Risk ratio; TTM: Target temperature management.

**Supplementary Information**
The online version contains supplementary material available at https://doi.org/10.1186/s13054-022-03935-z.

**Additional file 1.** Table 1. PRISMA checklist. Table 2. ROBUST criteria. Table 3. Sample search string. Table 4. Settings of the data driven priors for the primary outcome mortality and secondary outcome unfavourable neurology at 90-180 days. Table 5. GRADE assessment of certainty of evidence. Table 6. Sensitivity analyses using different effect size priors for the Bayesian meta-analysis. Figure 1. Frequentist meta-analysis of death and unfavourable neurological outcome. Figure 2. Sensitivity analysis for survival comparing initial shockable vs. non-shockable rhythms. Figure 3. Sensitivity analysis for neurological outcome comparing initial shockable vs. non-shockable rhythms.

**Authors’ contributions**
AA and SF contributed to study conception and design. The search and risk of bias assessment were performed independently by AA and MS and any discrepancies were resolved by consensus involving SF. Statistical analyses were performed by AA and SF. All authors participated in data interpretation and writing of the manuscript. The corresponding author attests that all listed authors meet authorship criteria and that no others meeting the criteria have been omitted. All authors read and approved the final manuscript.

**Funding**
Not applicable.

**Availability of data and materials**
The data are available upon reasonable request and published source data are already available in the public domain. No custom code was used and statistical packages are stated in Methods and available in the public domain.

**Declarations**
Ethics approval and consent to participate Not applicable.
Consent for publication
Not applicable.

Competing interests
MS has received speaker’s fees and travel grants from BARD Medical (Ireland). None of the other authors have any conflicts of interest to declare.

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Received: 6 October 2021 Accepted: 26 February 2022
Published online: 12 March 2022

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