The study of the validity and reliability of the Occupational Self-Assessment—traditional Chinese version

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Abstract

Introduction: Health professionals who work in mental healthcare settings need to use standardised, objective instruments; however, it is also extremely important that they maintain a client-focused perspective. The purpose of this study was to investigate the validity and reliability of the ‘Traditional Chinese version of the Occupational Self-Assessment’ (TC-OSA).

Methods: A total of 593 participants with mental illnesses participated in the study. The data were analysed using confirmatory factor analysis as well as the Rasch measurement model.

Results: The results of our analysis revealed that the TC-OSA encompasses four domains: self-performance (11 items), self-habituation (5 items), self-volition (5 items) and environment (8 items). Most of the items within each domain were found to have a good fit with the Rasch measurement model, whereas the CFA index was found to have a good fit for only three of the domains, the one exception being the environment domain.

Conclusions: We suggest applying the scale in clinical practice to identify the priority of intervention and as a measure for changes in outcomes. Further development and refinement of the environmental domain is guaranteed.

Keywords
Assessment, model of human occupation, client centre, confirmatory factor analysis, Rasch analysis

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Introduction

There is growing awareness that people are faced with greater stress and environmental pressures than ever before (American Psychological Association, 2017), which result in such people being increasingly vulnerable to stress-related diseases. Occupational therapists (OT), as important members of the psychiatric professional team, have a long history of providing psychosocial evaluations and programs to clients with psychiatric disorders; thus, the involvement of OT practitioners in the treatment of clients with mental illnesses could provide additional care perspectives in psychiatric healthcare delivery. Since clients with mental disorders, such as mood disorders or anxiety disorders, often differ from those with psychosis, in terms of both their characteristics and functional prognosis (Wang et al., 2003), the development of adequate assessment tools for the evaluation of clients with mental illnesses is crucial for OT practitioners.

An appropriate assessment tool for use in evaluating the strengths and limitations of clients is essential to the

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provision of quality care; furthermore, it is important that the instrument contents are both consistent with professional philosophy and applicable to the clients being served (Kielhofner et al., 2009). Since clients with mental disorders represent a group of people with a wide variety of problems, prognoses and functional outcomes, they need to be viewed from a holistic viewpoint, and should be empowered so that they can maintain control of their own lives. There were a few reviews of assessment tools for use in mental health occupational therapy practices (Hsiao et al., 2000); among those assessment being reviewed, we only adapted the comprehensive occupational therapy evaluation (Chiu et al., 2019; Wang et al., 2012), the Canadian Occupational Therapy Measure (Pan et al., 2007) and the kinetic house tree person drawing (Li et al., 2014) in Taiwan. We do have our own assessment tools developed in our own culture such as the Chu’s hand function test, attention test and activity of daily living test and Pan’s self-reported daily living test (Chu & Hsieh, 2004; Pan et al., 2018). However, none of the aforementioned tools were developed based on client-centred approach and using self-reported format.

Thus, suitable approaches to be adapted for this population are the ‘model of human occupation’ (MOHO; original developer Prof Keilhofner) (Taylor, 2017) and a client-centered methodology (Law et al., 2005). MOHO, which addresses issues of occupational functioning from a systemic view whilst also providing a paradigm for OT practitioners, proposes that volition, habituation, performance and environment systems are all extremely important elements for individuals to function competently in society.

The volition system is conceptualised as a collection of thoughts and feelings pertaining to an individual’s abilities and effectiveness; personal causation, interests and values are the three components of volition delineating a person’s sense of capacity and efficacy, their preferences and their everyday life patterns. The habituation system comprises of role scripts and habit maps which are rules and expectations that enable an individual to act automatically and fit in with environmental conditions. The performance system represents the capabilities that are fundamental to the skilled performance of an individual; these capabilities comprise of motor, process and communication skills. Finally, the environment system includes social and physical aspects that enable or constrain the adaptability of individuals in their daily lives (Taylor, 2017).

A client-centred approach is an additional theoretical thought process linked to occupational self-assessment (OSA) (Baron, 2006) which places the emphasis on consumer rights and active participation. When a client-centred approach is adopted within clinics, it has the effect of empowering the clients to choose their own goals and set up treatment priorities for themselves (Law et al., 2005).

OSA was developed from the ‘Self-Assessment of Occupational Functioning’, which was underpinned by the MOHO and client-centred approach. Although the OSA has been shown by researchers to have appropriate psychometric qualities (construct validity, sensitivity, reliability, ability to detect change and scale validity) (Kielhofner et al., 2009, 2010; Taylor et al., 2011), it was felt that it would be more appropriate to fit the 29 items into four domains, since this would make it compatible with the MOHO.

The OSA is a self-reported questionnaire which addresses self and environmental domains, each item is rated on both competence and value perspectives on a four-point rating scale. The items for each system within the MOHO comprise of self-performance (11 items), self-habituation (5 items), self-volition (5 items) and environment (8 items). The importance of translating OSA into traditional Chinese for use in Taiwan is to provide a client-centred and self-reported comprehensive assessment for use in treatment planning and progress evaluation. Furthermore, for the translated scales to be comparable and equally adapted as the original tool, the validation of the translated scales is important to ensure the proper psychometrics of the scale (Fisher et al., 1992). The primary aim of the study was to validate the translated Traditional Chinese version of the Occupational Self-Assessment (TC-OSA) (focus on the competence part) on a group of clients with mental illnesses in Taiwan.

The hypotheses of the study are listed below:

1. There was accepted internal consistency of the translated OSA subscales.
2. There were four unidimensional constructs for OSA-defining volition, habituation, performance and environment domains of MOHO by Rasch analysis and confirmatory factor analysis (CFA).
3. The level of item difficulties for each domain was adequate.
4. There was adequate item and person separation for each domain (at least above 2).

Methods

Translation

We obtained the translation permit and translated the OSA into traditional Chinese. A bilingual person subsequently carried out a reverse translation of the OSA forms, after which we examined the differences between the two versions and carried out appropriate changes to ensure the reconciliation of any differences between the two versions. The final version of the
OSA was determined after experts agreed on its readability, comprehensibility and congruence with the original version.

Participants and procedures

Our study sample, which was obtained from the Department of Psychiatry of a university affiliated hospital, comprised of a total of 593 patients suffering from mental disorders, their diagnoses include schizophrenia, affective disorder, anxiety disorder and ‘others’ (Table 1). The subjects were recruited from the inpatient ward with stable medication use and symptoms, and from outpatient clinics which they attended follow-ups by psychiatrists regularly. The study was approved by the Institutional Review Board of a related institution. We analysed a total of 232 males (39.1%) and 361 females (60.9%), with a mean age of 37.79 (SD = 15.03). The sample for analysis was based on the self-reported TC-OSA and data on personal characteristics (such as gender and age).

Traditional Chinese version of Occupational Self-Assessment

The TC-OSA questionnaire comprises of 29 items scored on a four-point Likert scale, indicating the level of competence in the performance of an activity. The scale ranges from 1 (no discernible problem) to 4 (serious problem), with the 29 items being organised into four domains comprising of self-performance (11 items), self-habituation (5 items), self-volition (5 items) and environment (8 items). These four domains are designed to evaluate personal occupational capabilities (for example, ‘physically doing what I need to do’) and environmental resources (such as ‘the things I need to be productive’).

Data analysis

In order to determine both the reliability and validity of the TC-OSA scale for people with mental disorders, we carried out an evaluation of the internal consistency of the scale by calculating the Cronbach’s coefficient alpha for the total scale scores (Cronbach, 1951). The data sample was subsequently examined based upon ‘confirmatory factor analysis’ (CFA) using ‘structural equation modelling’ (SEM) and the LISREL statistical program (Jöreskog & Sörbom, 2017). The ‘comparative fit index’ (CFI > 0.9), ‘goodness-of-fit index’ (GFI > 0.9), Adjusted goodness of fit index (AGFI > 0.9) and the ‘root mean square error of approximation’ (RMSEA ≤ 0.08) were used for the criteria (Bentler, 1990; Steiger, 1990). The confirmatory factor analytic models were evaluated using the ‘maximum likelihood’ (ML) estimation, an approach known to be robust to departure from normality. Given that our approach in this study was to confirm the hypothesised four domains of the TC-OSA scale: self-performance, self-habituation, self-volition and environment, our research methodology involves the evaluation of a total of four models.

We employed the WINSTEPS software (Linacre, 2005) to carry out a Rasch model analysis (Rasch, 1960), with this model being used to estimate both item and person fit on a common scale. The appropriateness of the rating scale for each item was examined according to the following criteria: (Chen et al., 2015) (1) at least 10 observations in each category and a regular observation distribution across categories, (2) monotonically increasing average measures across categories, (3) monotonically increasing step calibrations and (4) category outfit mean square (MnSq) values <2.0.

Results

The demographic data of the participants are presented in Table 1 (as discussed above). The results on the reliability for each domain and the overall TC-OSA are reported in Table 2. The Cronbach’s alpha coefficients in this study were found to be 0.853 for the self-performance domain, 0.823 for the self-habituation domain, 0.856 for the self-volition domain and 0.811 for the environment domain, whilst the Cronbach’s alpha for the overall TC-OSA scale was 0.934. Thus, hypothesis one is supported. These results clearly indicate that the scale items have good internal consistency. The item-total correlation coefficients are also presented in Table 2, with the results showing that there were moderate to high correlations between the items (that is, 0.36 or above).

We carried out tests on the goodness-of-fit for the four different models and the empirical outcomes.
The results are summarised in Table 3, which shows that the environmental domain did not fit, whilst self-performance, self-habituation and self-volition did. Thus, we are willing to accept the domain of self-performance, self-habituation and self-volition as an unitary construct within the MOHO model.

Whilst the first model was rejected due to the large sample size ($\chi^2 = 2649.88$, df = 55, $p < 0.001$), the CFI, GFI, AGFI and RMSEA for this model barely met the criteria. The second model fit the criteria except for AGFI and RMSEA. The third model also fit the criteria. The final model of the TC-OSA did not fit with the criteria. The second hypothesis is partially supported by CFA.

The rating scale structure for all four models fit based on the suggested criteria. The results of the self-performance domain showed that the internal consistency values for items and persons as 0.97 and 0.83, these values were satisfactory; as shown in Table 4, all items in the self-performance domain fit well. The most difficult item was ‘Taking care of others for whom I am responsible’, whilst the easiest item was ‘Managing my basic needs’. The gender variable revealed differences between males and females in the ‘differential item functioning’ (DIF) for this domain. The scale can classify clients into seven levels of performance ability.

The results of the self-habituation domain showed that the internal consistency values for items (0.97) and persons (0.79) indicated satisfactory to very high internal consistency; as shown in Table 5, all items in the

### Table 2. Reliability of the TC-OSA scale.

| Item         | Cronbach’s Alpha | Item-total correlation range |
|--------------|------------------|-----------------------------|
| Self-performance 1–11 | 0.853            | 0.446–0.612                  |
| Self-habituation 12–16 | 0.823            | 0.564–0.648                  |
| Self-volition 17–21 | 0.856            | 0.576–0.709                  |
| Environment 22–29 | 0.811            | 0.435–0.604                  |
| TC-OSA scale 1–29 | 0.934            | 0.360–0.675                  |

TC-OSA: Traditional Chinese version of the Occupational Self-Assessment.

### Table 3. Comparative goodness-of-fit statistics for each domain.

| Fit indices | Recommended value | Self-performance (1–11) | Self-habituation (12–16) | Self-volition (17–21) | Environment (22–29) |
|-------------|-------------------|-------------------------|--------------------------|-----------------------|----------------------|
| $\chi^2$    | N/A               | 2694.88                 | 1332.17                  | 1636.75               | 2087.58              |
| df          | N/A               | 55                      | 28                       | 28                    | 28                   |
| CFI         | $\geq 0.90$       | 0.96                    | 0.96                     | 0.99                  | 0.71                 |
| GFI         | $\geq 0.90$       | 0.96                    | 0.96                     | 0.99                  | 0.97                 |
| AGFI        | $\geq 0.90$       | 0.89                    | 0.89                     | 0.98                  | 0.62                 |
| RMSEA       | $\leq 0.08$       | 0.121                   | 0.135                    | 0.047                 | 0.224                |

CFI: comparative fit index; GFI: goodness of fit index; AGFI: adjusted goodness of fit index; RMSEA: root mean square error of approximation.

### Table 4. Self-performance domain item statistics in item measure order.

| Item | Measure | S.E. | Infit | Outfit | DIF<sup>c,d</sup> |
|------|---------|------|-------|--------|-------------------|
| 8.   | –1.00   | 0.07 | 1.07  | 1.20   | 1.01              |
| 4.   | –0.25   | 0.06 | 0.90  | 1.05   | 0.87              |
| 6.   | –0.12   | 0.06 | 1.16  | 1.30   | 1.18              |
| 9.   | –0.02   | 0.06 | 0.95  | 0.92   | 1.02              |
| 3.   | 0.02    | 0.06 | 0.94  | 0.93   | 1.02              |
| 10.  | 0.09    | 0.06 | 0.91  | 0.90   | 0.88              |
| 7.   | 0.14    | 0.06 | 1.18  | 0.99   | 1.02              |
| 2.   | 0.17    | 0.06 | 0.96  | 0.94   | 1.09              |
| 11.  | 0.28    | 0.06 | 0.88  | 0.97   | 0.40              |
| 1.   | 0.32    | 0.06 | 1.01  | 0.91   | 0.90              |
| 5.   | 0.38    | 0.06 | 0.98  | –0.40  | 0.90              |

<sup>a</sup>Measure value refers to personal ability.
<sup>b</sup>Model S.E. refers to the standard error of the estimate.
<sup>c</sup>The ‘differential item functioning’ (DIF) refers to the difference between males and females.
<sup>d</sup>Significance of the Bonferroni multiple-comparison correction: $p = 0.05$ divided by the DIF value ($0.05/11 = 0.0045$).
The self-habituation domain fit well. The most difficult item was ‘Relaxing and enjoying myself’ whilst the easiest item was ‘Handling my responsibilities’. No DIF was discernible for this domain. The scale can classify clients into six levels of habitation capability.

The results of the self-volition domain showed that the internal consistency values for items (0.94) and persons (0.82) indicated satisfactory to very high internal consistency; as shown in Table 6, all items in the self-volition domain fit well. The most difficult item was ‘Working towards my goals’, whilst the easiest item was ‘Doing activities I like’. No DIF was discernible for this domain. The scale can classify clients into four levels of volitional status.

The results of the environment domain showed that the internal consistency values for items (0.99) and persons (0.80) indicated satisfactory to very high internal consistency; as shown in Table 7, all items in the environment domain fit well. The most difficult item was ‘A place to live and take care of myself’, whilst the easiest item was ‘The basic things I need to live and take care of myself’. The gender variable revealed DIF for this domain. The scale can classify clients into eight levels of environmental resources.

The second, third and fourth hypothesis are supported by Rasch analysis.

Discussion

The TC-OSA was developed as a mean of assessing the occupational competence and values for specific tasks undertaken by people with various needs, whilst our study demonstrates that the OSA is a four-domain scale encompassing self-volition, self-habituation, self-performance and environment. The results of the confirmatory factor analysis, internal consistency and item-total correlations provide evidence of unidimensionality of the self-performance, self-habituation, and self-volition domain from a classical test theory perspective. Nevertheless, since the results of the factor analysis could vary due to sampling differences, the results of the CFA should be adapted carefully (Wang, 1997). The Rasch analyses demonstrated that all four domains of the TC-OSA fitted well, with gender DIF found for self-performance (item 11) and environment (item 26). Since this study is one of several studies that applied translated OSA to a group of clients with mental illnesses in Taiwan and demonstrated
that there are acceptable cross-cultural psychometric qualities for TC-OSA, it means that OTs who work with Taiwanese clients can readily apply it to generate client-centred treatment goals and monitor their progress.

The TC-OSA has important implications in clinical settings since the order of item difficulty can be used as an index to guide treatment priorities. For example, if a client can take care of his or her own needs, then the next task to train would be ‘Getting where I need to go’. Another example would be if a client has no problem of having ‘People who support and encourage me’, then the therapist would move to identify ‘People who do things with me’. Our approach reveals that the occupational competence of an individual can be seen from the perspectives of volition, habituation, performance and environment. For example, a client with Schizophrenia may have low motivation toward activity participation (volition), but he or she might be able to compensate by establishing daily routine or skill competence which both contribute to competent functional performance. Thus, the TC-OSA offers a way to identify the strengths and weaknesses of our clients and provides a framework to target treatment.

Our results were different from previous literature (Kielhofner et al., 2010; Taylor et al., 2011) in that our study results support the four dimensions (domains) including volition, habituation, performance and environment, whilst the environmental domain will need further development. Therefore, OTs can adapt OSA scores to understand the volitional, habitational or performance status of the client. Our argument is that volition, habituation, performance and environment are related, but demonstrate distinct domains.

In conclusion, our study shows that the reliability and validity of the TC-OSA are acceptable; thus, we suggest that therapists use the domain scores to represent the capabilities of their clients. However, the environment domain clearly needs to be modified to become an independent construct. The limitations of our study include the limited variety of the diagnostic group and the absence of measures of the cognitive competence of the participants who completed the self-reported questionnaire. There was also a lack of cross-cultural study of test–retest reliability, criterion-related validity and minimally detectable changes. In the future, there are needs to conduct research which include a more diverse sample with different capability level to demonstrate the sensitivity of the TC-OSA. Furthermore, it is necessary to modify the environmental scale so that the OT can apply it to allocate resources for our clients.

**Contributorship**

The first author drafted the article, analysed the data and finalised the article. The second author analysed the data and approved the results. The third author designed the acquisition of the data and contributed to the idea of the research. The fourth author assisted to acquire the data and collaborated on the data analysis and results interpretation.

**Ethical approval**

The data source was from several studies listed below. NTUH 9461701109 2005, NTUH 200804025R 2008, NTUH: National Taiwan University Hospital.

**Declaration of conflicting interests**

The author(s) declared no potential conflicts of interest with respect to the research, authorship, and/or publication of this article.

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**Table 7.** Environment domain item statistics in item measure order.

| Item                                                                 | Measure | S.E. | MnSq  | ZSTD  | MnSq  | ZSTD  | DIF  |
|----------------------------------------------------------------------|---------|------|-------|-------|-------|-------|------|
| 22. A place to live and take care of myself                         | -0.58   | 0.06 | 1.12  | 2.10  | 1.11  | 1.90  | 0.04 |
| 26. People who support and encourage me                             | -0.18   | 0.06 | 1.03  | 0.70  | 1.05  | 0.90  | 0.34*|
| 28. Opportunities to do things that I value and like                 | 0.16    | 0.06 | 0.97  | -0.50 | 0.98  | -0.40 | 0.07 |
| 25. The things I need to be productive                              | 0.27    | 0.06 | 1.04  | 0.70  | 1.06  | 1.10  | -0.28|
| 23. A place where I can be productive (work, study and volunteer)   | 0.33    | 0.06 | 1.03  | 0.60  | 1.05  | 1.00  | -0.26|
| 27. People who do things with me                                    | 0.36    | 0.06 | 0.87  | -2.50 | 0.87  | -2.60 | 0.13 |
| 29. Places where I can go and enjoy myself                          | 0.51    | 0.06 | 0.93  | -1.50 | 0.92  | -1.60 | 0.00 |
| 24. The basic things I need to live and take care of myself         | 0.87    | 0.07 | 0.98  | -0.30 | 0.97  | -0.60 | 0.02 |

MnSq: mean square; ZSTD: z-standardized.

*Measure value refers to personal ability.

Model S.E. refers to the standard error of the estimate.

The ‘differential item functioning’ (DIF) refers to the difference between males and females.

Significance of the Bonferroni multiple-comparison correction: $p = 0.05$ divided by the DIF value ($0.05/8 = 0.0063$).
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References
American Psychological Association. (2017, November). Stress in America: The state of our nation. Stress in America Surveys [Press release], http://www.apa.org/news/press/releases/stress/2017/state-nation.pdf
Baron, K. (2006). A user’s manual of occupational self assessment Version 2.2: Model of human occupation clearinghouse. The University of Illinois at Chicago, Chicago, Illinois.
Bentler, P. M. (1990). Comparative fit indexes in structural models. Psychological Bulletin, 107(2), 238–246.
Chen, Y. L., Pan, A. W., Chung, L., & Chen, T. J. (2015). Examining the validity and reliability of the Taita Symptom Checklist using Rasch analysis. Journal of the Formosan Medical Association, 114, 221–230.
Chiu, E. C., Lai, K. Y., Lin, S. K., Tang, S. F., Lee, S. C., & Hsieh, C. L. (2019). Construct Validity and Reliability of the Comprehensive Occupational Therapy Evaluation Scale (COTES) in people with schizophrenia. The American Journal of Occupational Therapy, 73, 7306205060. https://doi.org/10.5014/ajot.2019.026807
Chu, T.-F., & Hsieh, C.-L. (2004). Reliability and validity of the activities of daily living rating scale III for psychiatric patients: A preliminary study. Journal of Occupational Therapy Association R.O.C., 22, 1–14.
Cronbach, L. (1951). Coefficient alpha and the internal structure of tests. Psychometrika, 16(3), 297–334.
Fisher, A. G., Liu, Y., Velozo, C. A., & Pan, A. W. (1992). Cross-cultural assessment of process skills. American Journal of Occupational Therapy, 46(10), 876–885.
Hsiao, S., Pan, A., Chung, L., & Lu, S. (2000). The use of assessment tools in mental health OT: A national survey. Occupational Therapy Journal of OT Association of R.O. C., 18, 19–32.
Jöreskog, K. G., & Sörbom D. (2017). LISREL9.3: A guide to the program and applications (2nd ed.). SPSS Inc.
Kielhofner, G., Dobria, L., Forsyth, K., & Kramer, J. (2010). The occupational self assessment: Stability and the ability to detect change over time. OTJR: Occupation, Participation and Health, 30(1), 11–19.
Kielhofner, G., Forsyth, K., Kramer, J., & Iyenger, A. (2009). Developing the occupational self assessment: The use of Rasch analysis to assure internal validity, sensitivity and reliability. British Journal of Occupational Therapy, 72(3), 94–104.
Law, M. C., Baptiste, S., Carswell, A., McColl, M. A., Polatajko, H. J., & Pollock, N. (2005). Canadian occupational performance measure (4th ed.). CAOT.
Li, C. Y., Pan, A. W., Chung, L., Hsiung, P. C., Chen, T. J., & Liu, S. K. (2014). Psychometric study of the KHTTP scoring system for people with psychiatric disorders. Hong Kong Journal of Occupational Therapy, 24, 20–27.
Linacre, J. (2005). A user’s guide to winsteps/ministeps Rasch-model programs. MESA Press.
Pan, A. W., Chiao, H. Y., Hsiung, P. C., Chen, T. J., & Chung, L. (2007). The validity study of the Canadian Occupational Performance Measure – Chinese version in persons with schizophrenia in Taiwan. Journal of Taiwan Occupational Therapy Association, 25(2), 1–14.
Pan, A. W., Wu, C. Y., Chung, L., & Chen, T. J. (2018). Reliability and validity of the self-reported activities of daily living scale for people with mental illness. Hong Kong Journal of Occupational Therapy, 31(2), 115–124.
Rasch, G. (1960). Probabilistic models for some intelligence and attainment tests. University of Chicago Press.
Steiger, J. H. (1990). Structural model evaluation and modification: An interval estimation approach. Multivariate Behavioral Research, 25(2), 173–180.
Taylor, R. (2017). Kielhofner’s Model of human occupation, 5th edition: Theory and application. Lippincott Williams & Wilkins.
Taylor, R., Lee, S. W., Kramer, J., Shirashi, Y., & Kielhofner, G. (2011). Psychometric study of the Occupational Self Assessment with adolescents after infectious mononucleosis. American Journal of Occupational Therapy, 65(2), e20–e28.
Wang, S. M., Pan, A. W., Liu, L., & Chung, L. (2003, June). The application of the occupational self-assessment in Taiwan: Study of the psychometric qualities. [Paper presentation] The 3rd Asian-Pacific Occupational Therapy Congress, Singapore.
Wang, W. C. (1997). The construct of the test: Factor analysis or Rasch analysis. Survey Research, 3, 129–166.
Wang, Y. T., Pan, A. W., Chung, L., Chen, T. J., Liu, L. T., & Chen, Y. L. (2012). The use of confirmatory factor analysis and Rasch analysis to examine the psychometric qualities of the comprehensive occupational therapy evaluation scale in persons with mental illness. Formosa Journal of Medicine, 16, 212–218.