Influence of hospital volume on nephrectomy mortality and complications: a systematic review and meta-analysis stratified by surgical type

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ABSTRACT

Objectives The provision of complex surgery is increasingly centralised to high-volume (HV) specialist hospitals. Evidence to support nephrectomy centralisation however has been inconsistent. We conducted a systematic review and meta-analysis to determine the association between hospital case volumes and perioperative outcomes in radical nephrectomy, partial nephrectomy and nephrectomy with venous thrombectomy.

Methods Medline, Embase and the Cochrane Library were searched for relevant studies published between 1990 and 2016. Pooled effect estimates for nephrectomy mortality and complications were calculated for each nephrectomy type using the DerSimonian and Laird random-effects model. Sensitivity analyses were performed to examine the effects of heterogeneity on the pooled effect estimates by excluding studies with the heaviest weighting, lowest methodological score and most likely to introduce bias from misclassification of standardised hospital volume.

Results Some 226 372 patients from 16 publications were included in our review and meta-analysis. Considerable between-study heterogeneity was noted and only a few reported volume–outcome relationships specifically in partial nephrectomy or nephrectomy with venous thrombectomy. HV hospitals were correlated with a 26% and 52% reduction in mortality for radical nephrectomy (OR 0.74, 95% CI 0.61 to 0.90, p<0.01) and nephrectomy with venous thrombectomy (OR 0.48, 95% CI 0.29 to 0.81, p<0.01), respectively. In addition, radical nephrectomy in HV hospitals was associated with an 18% reduction in complications (OR 0.82, 95% CI 0.73 to 0.92, p<0.01). No significant volume–outcome relationship in mortality (OR 0.84, 95% CI 0.31 to 2.26, p=0.73) or complications (OR 0.85, 95% CI 0.55 to 1.30, p=0.44) was observed for partial nephrectomy.

Conclusions Our findings suggest that patients undergoing radical nephrectomy have improved outcomes when treated by HV hospitals. Evidence of this in partial nephrectomy and nephrectomy with venous thrombectomy is however not yet clear and could be secondary to the low number of studies included and the small patient number in our analyses. Further investigation is warranted to establish the full potential of nephrectomy centralisation particularly as existing evidence is of low quality with significant heterogeneity.

INTRODUCTION

In recent years, there has been an emerging trend for the centralisation of complex operations in healthcare systems around the world.1–3 This shift is supported by the growing research and evidence suggesting that hospitals and surgeons with high caseloads have better patient outcomes.4–8 Proponents argue that centralisation allows more effective use of clinical expertise and specialist equipment, and the increased exposure improves surgical skills and provides better training opportunities. Centralisation can also facilitate quicker adoption of care pathways, such as enhanced recovery, and may have more long-term financial sustainability for hospitals. However, surgical centralisation requires further travel distance and limits patient choice when many would prefer to undergo surgery locally even if greater mortality risks are taken into consideration.9 Differences in
Renal cancer accounts for over 2% of all new cancer diagnoses worldwide affecting more than 330,000 individuals annually. Widespread use of cross-sectional imaging and increasing prevalence of obesity have contributed to a rising renal cancer incidence in many countries. Despite recent developments in systemic therapies, nephrectomy is often considered the only potentially curative treatment for renal cancer, and the number of nephrectomies being performed is likely to increase as a result. It is, therefore, critical that health service providers understand the effects that organisational changes may have on patient outcomes. While there has been an expansion of volume–outcome research, no consensus has so far been reached on the efficacy of centralising nephrectomy, and many uncertainties remain about its potential benefits particularly as radical and partial nephrectomy carry different surgical complexities and outcomes. We present a contemporary systematic review and meta-analysis of the published literature on the association between hospital case volumes and perioperative outcomes stratified by nephrectomy types. We hypothesise that outcomes significantly improve with higher nephrectomy case volumes.

**METHODS**

**Search criteria and data extraction**

The systematic review and meta-analysis was reported in accordance to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines (online supplementary appendix 1). Medical subject heading terms and key words for nephrectomy, case volume and outcomes were used in Medline, Embase and the Cochrane Library to search for relevant studies published between January 1990 and December 2016 (online supplementary appendix 2). Studies published prior to 1990 were not considered as recent medical and surgical advancements would have limited their applicability to the modern healthcare system. Only studies published in English were considered as the risk of potential language bias associated with this exclusion generally has little effect on summary effect estimates. References were searched manually for additional relevant studies.

We included studies that presented original data in full-texts on adult nephrectomy outcomes across two or more hospital case volume categories. Abstracts, case reports and review articles were excluded (online supplementary appendix 3). No restriction was set on the study design and both prospective and retrospective studies were considered. Only those describing the volume–outcome relationships in radical nephrectomy with or without venous thrombectomy and/or partial nephrectomy were eligible. Paediatric cohorts were excluded as were articles comprised solely of nephroureterectomy or nephrectomy for non-oncological indications. Restriction on the reported outcomes was only applied at the end stage of the search to enable assessment of the current published evidence. Only studies reporting nephrectomy mortality and complications were included in the final analysis.

Two investigators (RCJH and JM) independently reviewed all studies for inclusion, data extraction and methodological quality. Any disagreement between the two reviewers was resolved by discussion and consultation with a third reviewer (JNA). Where only rates of outcomes were presented, these were applied to the case number to give the number of events, within the error of the published results. Study authors were contacted for further clarification if specific rates of outcomes and case numbers were not published.

As the cut-off values for hospital case volume categories differed among studies, we used the approach adopted by similar previous meta-analyses by dichotomising the volume groups presented by each study into low volume (LV) and high volume (HV) when the article presented an even number of volume groupings. If a study presented an odd number of volume groups, the middle group was considered as LV.

Methodological quality and potential risk of bias were scored using a 10-domain system designed to measure the degree in which the study is likely to reveal generalisable conclusion about the magnitude and nature of the volume–outcome relationship. Each domain provides a score between 0 and 3 with a total maximum of 18, suggesting a well-designed study. The parameters included the representativeness of the sample, the number of hospitals analysed, the samples size, the number of adverse events recorded, the appropriateness of patient selection, the number of volume categories examined, the number of outcomes measured, the degree of risk adjustment performed, whether hospital and surgeon case volumes were analysed in conjunction and whether clinical processes of care were measured.

If studies extracted data from the same source with overlaps in the study periods, we employed the following rules to avoid duplicating populations: (1) studies with identical patient cohort but examining different outcomes were considered and analysed separately, (2) studies that derived data from older datasets were excluded in favour of the more contemporary cohort, (3) if the above rules were not applicable, studies with the lower methodological quality scores were excluded and (4) where quality scores were equal, the study covering the longest period was included.

**Quantitative data synthesis**

All statistical analyses were performed using Stata 14. Nephrectomy types were categorised into radical nephrectomy, partial nephrectomy and nephrectomy with venous thrombectomy and analysed separately. Studies involving multiple types of nephrectomies were analysed based on the aforementioned groups, but if this was not feasible, they were categorised as radical nephrectomy. With the
assumption that a distribution of effects exists among studies, all pooled effect size were calculated using the DerSimonian and Laird random-effects model, which provided more conservative estimates compared with fixed-effect model. OR and 95% CI were calculated and presented for each outcome measure using LV groups as the reference.

When the meta-analysis demonstrated significantly better outcomes in HV hospitals, we quantified the clinical effectiveness of centralisation by calculating the numbers needed to treat, or in our case numbers needed to centralise (NNC). NNC represents the number of cases that will need to be centralised from LV hospitals and treated by HV hospitals in order to prevent one adverse event.

Heterogeneity
As the DerSimonian and Laird model would have only accounted for some between-study heterogeneity, we further quantified heterogeneity by calculating $I^2$ statistic. $I^2$ provides an easily understood number, which describes the proportion of total variation in estimates that is due to heterogeneity rather than chance. Values of 25% or lower denote low heterogeneity and values of 75% or greater denote considerable heterogeneity. Meta-regression was performed to explore the influence of potential explanatory variables on heterogeneity including each study’s publication year, country, data source, number of patients and their demographics, number of hospitals and threshold for HV hospitals.

Publication bias
Funnel plots were generated to investigate potential publication bias, and were enhanced to include contours that divide the funnel into statistically significant and non-significant areas. Funnel plot symmetry suggests low probability of publication bias and Harbord’s modified test was used to test for asymmetry. Harbord’s test reduces false positive rates when applied to binary outcome data, especially when there is low between-study heterogeneity.

Trim and fill method was also performed to account for publication bias by adjusting the meta-analysis to incorporate the theoretically missing studies.

Sensitivity analysis
To examine specific studies’ effects on pooled effect size, sensitivity analyses were performed by excluding individual studies and repeating the meta-analyses. We examined the effects of studies with the heaviest weighting and studies with the lowest methodological quality score. As there is currently no consensus on what nephrectomy case volume is necessary to be considered as HV, we repeated our analyses by excluding studies whose standardised HV categories overlapped most significantly with the standardised LV categories in other studies to account for potential bias of misclassifying volume categories in our dichotomy.

As secondary analyses, we additionally repeated the meta-analysis three further times with different methods of dichotomising the volume groups to examine whether our initial estimates would remain consistent. The methods of dichotomising were (1) lowest volume categories and all others, (2) even dichotomy and when studies present an odd number of volume categories, the middle group was considered as HV and (3) highest volume categories and all others.

RESULTS
Study selection and characteristics
From the 5680 articles initially identified, 16 were included in the systematic review containing 226,372 patients from six countries (figure 1). For the meta-analysis, 11 studies with 201,506 patients examining radical nephrectomy were included while 4 studies of 23,617 patients and 2 studies of 1249 patients examining partial nephrectomy and nephrectomy with venous thrombectomy were included, respectively. Publication year ranged from 2002 to 2016, while cohort periods covered from 1993 to 2013.

Table 1 summarises the characteristics of the included studies. Variations were observed in study designs including source of data and outcomes measured. Out of a maximum possible score of 18, the median quality score from the included studies was nine (IQR 8–9) with the majority of the studies failing to adequately address potential confounders including measuring the appropriateness of patient selection, adjusting for case-mix variations and accounting for differences in clinical risks and processes of care (online supplementary appendix 4). Variable thresholds for HV hospitals were noted across the included studies.

Mortality and hospital volumes
Postoperative mortality, defined as inpatient or 30 day, was the most frequently examined outcome reported in 14 studies. Ten studies reported mortality in radical nephrectomy, three in partial nephrectomy and two in nephrectomy with venous thrombectomy. The overall mortality rate was 1.59% (range 0.20%–7.2%) with mortality rates in HV and LV hospitals being 1.47% and 1.68%, respectively.

Radical nephrectomy
Meta-analysis demonstrated that patients who underwent radical nephrectomy in HV hospitals had a 26% reduction in postoperative mortality, corresponding to an NNC of 234 (figure 2A).

Significant heterogeneity was observed ($I^2=75.0\%$, $p<0.01$). Meta-regression was performed to investigate the potential explanatory variables for heterogeneity, and only differences in the threshold values for HV hospitals were shown to be a significant contributor (online supplementary appendix 5A). Subgroup analysis of the three studies examining exclusively radical nephrectomies
demonstrated a more pronounced reduction in postoperative mortality favouring HV hospitals (OR 0.62, 95% CI 0.53 to 0.71, p<0.01). This corresponded to a lower NNC of 166 with little residual heterogeneity ($I^2=0.0\%$, p=0.40). The overall funnel plot was visually asymmetrical particularly missing studies with effect estimates favouring LV hospitals (figure 3). However, Harbord’s modified test did not show significant asymmetry (p=0.40) and ‘trim and fill’ method did not change the initial estimate, indicating no clear evidence of publication bias.

In sensitivity analyses, exclusion of the most heavily weighted study led to a similar pooled-effect estimate (OR 0.70, 95% CI 0.55 to 0.88, p<0.01). Exclusion of the study with the lowest quality score also did not significantly alter our result (OR 0.74, 95% CI 0.61 to 0.91, p<0.01). To examine the potential bias introduced by misclassification of hospital volume, two cohorts were excluded. This did not substantially change our pooled-effect estimate either (OR 0.73, 95% CI 0.58 to 0.93, p=0.01). Overall, radical nephrectomies in HV hospitals appeared to have significantly lower mortality.

Partial nephrectomy
Meta-analysis showed that partial nephrectomy patients operated in HV hospitals had a 16% reduction in postoperative mortality, but this was not statistically significant (OR 0.84, 95% CI 0.31 to 2.26, p=0.73) (figure 2B). Moderate but non-significant heterogeneity was noted ($I^2=36.84\%$, p=0.21). Sensitivity analyses removing studies with the heaviest weighting or most likely to introduce misclassification bias of exposure demonstrated reduced mortality favouring HV hospitals, but these remained non-significant.

Nephrectomy with venous thrombectomy
Patients who underwent nephrectomy with venous thrombectomy in HV hospitals had a 52% reduction (OR 0.48, 95% CI 0.29 to 0.81, p<0.01) in short-term mortality compared with LV hospitals (figure 2C). This corresponded to an NNC of 25 with low heterogeneity ($I^2=0.0\%$, p=0.50). Due to the small number of studies reporting this outcome, further testing of heterogeneity and publication bias was not expected to generate meaningful results and this was not attempted.

Complications and hospital volumes
Complication was the second most frequently investigated outcome, reported in 11 studies. Events considered as a complication differed among studies (table 1). Seven studies reported complications in radical
| Study characteristics | Outcomes measured |
|----------------------|-------------------|
| **Reference**        | **Year** | **Country** | **Period** | **Data type** | **No of patients** | **No of hospitals** | **Low-** | **High-** | **Mortality** | **Complications (breakdowns)** | **Transfusion** | **LOS** | **Conversions** | **Others** | **Case mix**                | **Quality score (18)** |
| ---------------------|----------|--------------|------------|---------------|-------------------|--------------------|----------|----------|--------------|-----------------------------|-----------------|--------|----------------|----------|-----------------------------|---------------------|
| Hjelle^10^           | 2016     | Norway       | 2008–2013  | Admin         | 3273             | 40                 | 20       | 40       | 30 day       | –                          | –                | –      | –              | +        | –                          | 8                   |
| Becker^11/ Sun^12^   | 2014/2012 | USA         | 1998–2007  | Admin         | 48172            | N/S                | 5        | 16       | In-hospital  | (+ 17 events including haemorrhage, cardiac arrest, infection, wound disruption, seroma, pneumothorax, VTE, etc) | +                | +      | –              | –        | Demographics, tumour stage, nephrectomy type | 9                   |
| Hanchanale^32^       | 2010     | England      | 1998–2005  | Admin         | 20672            | 1181               | 14       | 35       | In-hospital  | –                          | –                | –      | +              | –        | Demographics                  | 9                   |
| Yasunaga^33^         | 2010     | Japan        | 2006–2007  | Admin         | 7988             | 646                | 26       | 65       | In-hospital  | (+ 11 events including surgical site infection, UTI, VTE, sepsis, ileus, stroke, cardiac events, renal, failure, peritonitis, etc) | –                | –      | –              | –        | Demographics, comorbidity, nephrectomy type, laparoscopy, payer/hospital type | 9                   |
| Mitchell^34^         | 2009     | USA          | 2003–2007  | Clinical      | 42988            | 134                | 99/4.5 year | 500/4.5 year | In-hospital  | (+ not specified) | –                | +      | –              | –        | ICU admission                  | None                | 9                   |
| Yasunaga^35^         | 2008     | Japan        | 2006–2007  | Clinical      | 1704             | 461                | 9        | 40       | In-hospital  | (+ wound infection, pneumonia, ileus, renal dysfunction, others) | –                | –      | –              | OT, EBL  | Demographics, comorbidity, laparoscopy, hospital type, tumour location | 11                  |
| Davenport^36^        | 2005     | England      | 2004       | Clinical      | 598              | 48                 | <1/month | >1/month | N/S          | (+ 12 events including bleeding, bowel injury, GI bleed, renal failure, pneumothorax, VTE, MI, splenic injury, etc) | +                | –      | +              | OT       | None                        | 4                   |
| Keoghane^37^         | 2004     | England      | 2001–2002  | Clinical      | 263              | 25                 | 5        | 6        | –            | (+ 16 events including renal failure, sepsis, wound infection, bowel injury, incisional hernia, perihepatic collection, etc) | –                | –      | +              | –        | None                        | 3                   |
| Taub^38^             | 2004     | USA          | 1993–1997  | Admin         | 16858            | 962                | 14       | 34       | In-hospital  | –                          | –                | –      | +              | –        | Demographics, comorbidity, admission acuity | 9                   |
### Table 1 Continued

| Study characteristics | Outcomes measured | Quality score (18) |
|-----------------------|-------------------|-------------------|
| **Partial nephrectomy** |                  |                   |
| Birkmeyer et al. 2002 USA 1994-1999 Admin 58990 3292 6 33 30 day or in-hospital - - - - | Demographics, comorbidity, admission acuity 7 |
| Couapel 2014 France 2010 Clinical 570 53 4/7 month 19/7 month N/S + (medical and surgical events, not further specified) - + + OT, EBL, totalisation, +ve margin N/S | 8 |
| Monn 2014 USA 2009-2011 Admin 17583 322 13 35 - + (organ-based complications not further specified, pain, seroma, shock, haematoma, hypotension, VTE, pneumothorax) + + - Hospital cost | 9 |
| Abouassaly 2012 Canada 1998-2008 Admin 4292 181 146/10 year 797/10 year In-hospital + (not specified) - - - - | Demographics, comorbidity, payer, region, hospital type 12 |
| Taub 2004 USA 1993-1997 Admin 1172 962 14 34 In-hospital - - + - | Demographics, comorbidity, admission acuity 9 |
| **Nephrectomy with venous thrombectomy** |                  |                   |
| Toren 2013 Canada 1998-2008 Admin 816 120 N/S N/S In-hospital + (40 medical and surgical events including MI, CHF, PE, infection, organ injury, pneumothorax etc) + - - - | Demographics, comorbidity, region 11 |
| Yap 2012 Canada 1995-2004 Admin 433 N/S 2/10 year 8/10 year 30 day - - - - | Demographics 11 |

*Low, cut-off value for lowest volume group; high, cut-off value for highest-volume group. Volume units are cases per year unless specified.
†Becker and Sun were equal in the data source used, outcomes evaluated, periods covered and quality scores, but employed different analyses to evaluate the benefit of regionalisation and volume-outcome relationship in nephrectomy.†33 We therefore treated them as one single cohort, with no duplicates in our analysis.

Admin, administrative; CHF, congestive heart failure; EBL, estimated blood loss; GI, gastrointestinal; ICU, intensive care unit; LOS, length of stay; MI, myocardial infarction; N/S, not specified; OT, operating time; PE, pulmonary embolism; UTI, urinary tract infection; VTE, venous thromboembolism.
nephrectomy 18 30 31 33–35 41 and three in partial nephrec-
tomy.38 39 42 Only one study examined volume–outcome
relationship in nephrectomy with venous thrombectomy
and meta-analysis was, therefore, not appropriate.40 The
overall complication rate was 16.26% (range 7.4%–78%).
HV hospitals had complication rates of 15% compared
with 17.51% in LV hospitals.

Radical nephrectomy
Meta-analysis showed an 18% reduction (OR 0.82, 95% CI
0.73 to 0.92, p<0.01) in nephrectomy complications in
HV centres, corresponding to an NNC of 38 (figure 4A).
Significant heterogeneity was noted (I²=76.25%, p<0.01),
but none of the factors examined in meta-regression
significantly contributed to this (online supplementary
appendix 5B). Sensitivity analyses by removing studies
with the lowest quality41 or most likely to introduce
misclassification bias of exposure35 41 did not signifi-
cantly alter our initial result. Excluding study with the
heaviest weighting however led to a loss of significance
in the pooled-effect estimate, which however still demon-
strated a 11% reduction in complications in HV hospitals
(OR 0.89, 95% CI 0.74 to 1.08, p=0.24). Overall, radical
nephrectomies performed in HV hospitals appeared to
have significantly lower complications compared with LV
hospitals.

Partial nephrectomy
Partial nephrectomy patients operated in HV hospitals
had a 15% reduction in complications, but this was not
statistically significant (OR 0.85, 95% CI 0.55 to 1.30,
p=0.44) (figure 4B). Significant heterogeneity was noted
(I²=94.80%, p<0.01). Sensitivity analysis by removing
studies with the heaviest weighting42 or most likely to
introduce misclassification bias of hospital volume38 did
not result in significance.

Secondary analyses using different methods for
dichotomising HV and LV
As there is no consensus on what constituted HV hospi-
tals in current evidence, simple dichotomy of volume
groups may introduce inherent bias to the estimates. Yet
no recommendation on how best to proceed in volume–
outcome analysis presently exists.
In our secondary analyses, we consistently observed significantly lower risks of mortality for both radical nephrectomy and nephrectomy with venous thrombectomy in HV hospitals regardless how volumes were dichotomised (table 2). The magnitudes of risk reductions were more pronounced when higher thresholds for HV hospitals were considered particularly for radical nephrectomy mortality. Partial nephrectomy mortality however continued to demonstrate no significant association to volume even when dichotomies were comparing the highest volume groups to all others.

Risks of radical nephrectomy complications remained significantly reduced in HV hospitals when the dichotomy threshold for HV hospitals was increased, but the significance was lost when the thresholds were lowered. Association between hospital volumes and partial nephrectomy complications remained insignificant regardless of how HV was defined in our dichotomy.

**DISCUSSION**

Evidence on volume–outcome relationships in complex diseases and procedures has increased substantially in recent years. Many operations have been shown to have improved outcomes in HV centres, but this may not be uniform across all surgeries and benefits have not been
associated with volume in percutaneous nephrolithotomy or appendicectomy.44–48 This meta-analysis provides a contemporary review of the effects of centralisation in nephrectomy outcomes. It reveals significant inverse associations between hospital case volumes for short-term mortality and complications for radical nephrectomy, but evidence of these for partial nephrectomy and nephrectomy with venous thrombectomy remains less compelling.

Considered individually, all but three studies in our review reported lack of associations between hospital volume and nephrectomy mortality.29 36 37 However, such associations in favour of HV hospitals were apparent when considering the totality of the evidence particularly in radical nephrectomy and venous thrombectomy. This finding is consistent with the only other meta-analysis on nephrectomy volume–outcome relationship published in 2009 but includes fourfold greater number of studies.23 Our meta-analysis demonstrates that the mortality benefit seen in radical nephrectomy may be relatively small requiring centralisation of 234 patients in order to avoid one death. However, the NNC decreased considerably to 166 in our sensitivity analyses. Coupled with the much lower NNC of 38 for radical nephrectomy complications, there is moderate evidence to support its centralisation.

In our analyses, the mortality reduction for venous thrombectomy was observed to be more pronounced than that in radical nephrectomy. This is consistent with the ‘practice-makes-perfect’ hypothesis particularly as venous thrombectomy is a technically more challenging procedure compared with radical nephrectomy, though interestingly a similar trend was not observed for partial nephrectomy.49 These results should however be interpreted taking into consideration that only a few studies have so far reported on the volume–outcome relationships for partial nephrectomy and venous thrombectomy and the pooled effect estimates were synthesised from just two to three publications, thus the overall evidence is weak. As partial nephrectomy has only been widely adopted in the last two decades and nephrectomy centralisation also a relatively recent phenomenon, it is likely that more evidence will emerge in the coming years and repeating the meta-analysis at such point is warranted.1 50 This will be of particular importance as partial nephrectomy has been demonstrated to be a safe procedure and the relatively small number of partial nephrectomy patients in our meta-analysis might not have been sufficiently powered to reveal the true presence of a volume–outcome relationship, as evident in our wide CI.51 This may also explain the lack of significant association between partial nephrectomy complications and hospital volumes. With its low mortality and morbidity rates, other outcome measures such as ischaemic time and negative surgical margins, are likely to be more appropriate quality markers in volume–outcome analysis, but these have so far been poorly evaluated.

Despite the strict inclusion criteria in our studies, we observed considerable heterogeneity, especially in the meta-analyses of nephrectomy complications. One explanation for this is the lack of standardised reporting of complications by individual studies. Harder endpoints as previously discussed could have overcome this. Other more objective outcomes including transfusion rate and length of stay were reported by four28–30 31 33 35 42 and seven studies30–32 34 36 42 52 in our systematic review, respectively, but they were not in adequate numbers to be stratified by nephrectomy types or in sufficiently detailed data to

Table 2 Results of meta-analysis by using different methods of dichotomising volume groups

|                          | Lowest volume group vs all others | Even dichotomy, middle group as HV | Even dichotomy, middle group as LV (primary analysis) | Highest volume group vs all others |
|--------------------------|----------------------------------|-----------------------------------|------------------------------------------------------|----------------------------------|
| **Mortality**            |                                  |                                   |                                                      |                                  |
| Radical nephrectomy      | OR: 0.82 95% CI 0.71 to 0.94,   | OR: 0.83 95% CI 0.72 to 0.95,     | OR: 0.74 95% CI 0.61 to 0.90, p<0.01                 | OR: 0.72 95% CI 0.61 to 0.85, p<0.01 |
|                          | p=0.01                            | p<0.01                            | p=0.01                                               | p=0.01                           |
| Partial nephrectomy      | OR: 0.48 95% CI 0.18 to 1.31,    | OR: 0.67 95% CI 0.17 to 2.75,     | OR: 0.84 95% CI 0.31 to 2.26, p=0.73                 | OR: 0.93 95% CI 0.31 to 2.77, p=0.90 |
|                          | p=0.15                            | p=0.58                            | p=0.73                                               | p=0.90                           |
| Nephrectomy with venous  | OR: 0.59 95% CI 0.35 to 0.99,    | OR: 0.46 95% CI 0.27 to 0.80,     | OR: 0.48 95% CI 0.29 to 0.81, p=0.01                 | OR: 0.48 95% CI 0.25 to 0.92, p=0.03 |
| thrombectomy             | p=0.045                           | p<0.01                            | p=0.01                                               | p=0.03                           |
| **Complications**        |                                  |                                   |                                                      |                                  |
| Radical nephrectomy      | OR: 0.89 95% CI 0.78 to 1.01,    | OR: 0.84 95% CI 0.68 to 1.05,     | OR: 0.82 95% CI 0.73 to 0.92, p<0.01                 | OR: 0.82 95% CI 0.73 to 0.92, p<0.01 |
|                          | p=0.07                            | p=0.13                            | p=0.01                                               | p=0.01                           |
| Partial nephrectomy      | OR: 0.82 95% CI 0.55 to 1.41,    | OR: 0.80 95% CI 0.47 to 1.36,     | OR: 0.85 95% CI 0.55 to 1.30, p=0.44                 | OR: 0.81 95% CI 0.53 to 1.24, p=0.33 |
|                          | p=0.60                            | p=0.40                            | p=0.44                                               | p=0.33                           |

HV, high volume; LV, low volume.
perform meta-analyses. In addition, variations in the threshold values for HV hospitals likely contributed to the heterogeneity, although this was not evident in the meta-regression. Volume was also used as a proxy marker for surgical and care quality, but the precise clinical processes that may improve patient outcomes were not directly measured or identified. It is therefore conceivable for some heterogeneity to arise from these unmeasured practices. Results from the multiple sensitivity analyses to adjust for study differences have however remained robust and our study would still appear to be informative and relevant. There are other research designs that may be more appropriate in testing our hypothesis such as analysis of primary data amalgamated from multiple population cohorts. The considerable ethical concerns and logistical constraints of this may however be challenging to overcome and not practically feasible.

While there has been an expansion in the studies on nephrectomy volume–outcome relationship, many questions continue to be unanswered. The proportion of nephrectomy performed under laparoscopy or robotic assistance is growing.53 54 There is however a paucity of evidence specifically investigating this in the volume–outcome context with only one study examining the differences in perioperative measures in robotic partial nephrectomy.42 Three other studies have adjusted surgical techniques in multivariable regressions, but these did not directly demonstrate the effect of laparoscopic volumes on surgical outcomes.18 30 33 Due to the small study number and data quality, it was not possible in this meta-analysis to further stratify each nephrectomy type into open and minimally invasive and our results should be interpreted taking this limitation into account.

Tumour characteristics including TNM stage and grades are well established to significantly affect and predict nephrectomy mortality, but only two studies have so far adjusted for this in their analyses.18 29 Surgeon case volume and degree of specialisation also play significant roles in determining operative outcome, and can be more important than hospital case volume alone.45 55–57 While not the focus of this study, no significant association was found between surgeon volume and complications in radical nephrectomy, but 31% and 16% reduction in mortality and complications respectively was observed in partial nephrectomy in HV surgeons.39 HV surgeons performing nephrectomy with venous thrombectomy were also reported to have reduced risk of mortality,19 but this was not observed in a subsequent study.46 As our analyses were based on crude pooled effect estimates, future meta-analysis should ideally attempt to adjust for other possible confounders including patient demographics, socioeconomic status and comorbidities, although this may be methodologically challenging. It would be of high interest to understand the interactions among patient characteristics, surgeon volume, surgical approach and oncological factors in the volume–outcome relationship and may provide additional insights to selecting patients that will benefit the most from nephrectomy centralisation, such as those with multiple comorbidities or advanced disease. Similarly, no study has examined the long-term benefits of centralising nephrectomy when HV centres have been demonstrated to increase oncological survival in other cancer surgery.58 59 Results of this may further influence the recommendations for nephrectomy centralisation and this is currently being explored in our ongoing work. Other outcomes including long-term risks of chronic kidney disease and cardiovascular morbidities and patient reported outcome measures may also provide more relevant and holistic measurements of the potential efficacy of nephrectomy centralisation.

Our secondary analyses would suggest that a minimum volume threshold for nephrectomy likely exist, and beyond that, risks of adverse outcomes may continue to decrease with further increase in volume. An important limitation of this however is that this minimum threshold cannot be objectively determined from the current evidence. The specific care processes that may produce good outcomes, such as access to nurse specialists and clinical trials, could not be determined from our study. As volume is likely to be a proxy marker for quality, increasing volume alone in itself is unlikely to reduce adverse results.60 Future research should concentrate on identifying the qualitative differences between providers in order for the contributing good practices to be adopted by lesser performing centres.

**Conclusions**

Current evidence of the association between hospital volumes and nephrectomy outcomes is of low quality with considerable between-study heterogeneity. Our meta-analyses demonstrated significant reductions in mortality and complications for patients undergoing radical nephrectomy in HV hospitals. Evidence of this in partial nephrectomy and nephrectomy with venous thrombectomy is not yet clear but warrants further investigations.

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