Lower Bound on the Capacity of Continuous-Time Wiener Phase Noise Channels

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Abstract—A continuous-time Wiener phase noise channel with an integrate-and-dump multi-sample receiver is studied. A lower bound to the capacity with an average input power constraint is derived, and a high signal-to-noise ratio (SNR) analysis is performed. The capacity pre-log depends on the oversampling factor, and amplitude and phase modulation do not equally contribute to capacity at high SNR.

I. INTRODUCTION

Instabilities of the oscillators used for up- and down-conversion of signals in communication systems give rise to the phenomenon known as phase noise. The impairment on the system performance can be severe even for high-quality oscillators, if the continuous-time waveform is processed by long filters at the receiver side. This is the case, for example, when the symbol time is very long, as happens when using orthogonal frequency division multiplexing.

Typically, the phase noise generated by oscillators is a random process with memory, and this makes the analysis of the capacity challenging. The phase noise is usually modeled as a Wiener process, as it turns out to be accurate in describing the capacity challenging. The phase noise is usually modeled

\[ Y(t) = X(t)e^{j\Theta(t)} + W(t), \quad 0 \leq t \leq T \]

where \( j = \sqrt{-1} \), \( X(t) \) is the data bearing input waveform, and \( W(t) \) is a circularly symmetric complex white Gaussian noise. The phase process is given by

\[ \Theta(t) = \Theta(0) + \gamma \sqrt{T} B(t/T), \quad 0 \leq t \leq T, \]

where \( B(\cdot) \) is a standard Wiener process, i.e., a process characterized by the following properties:

- \( B(0) = 0 \),
- for any \( 1 \geq t > s \geq 0 \), \( B(t) - B(s) \sim \mathcal{N}(0, t-s) \) is independent of the sigma algebra generated by \( \{ B(u) : u \leq s \} \),
- \( B(\cdot) \) has continuous sample paths.

One can think of the Wiener phase process as an accumulation of white noise:

\[ \Theta(t) = \Theta(0) + \gamma \int_0^t B'(\tau) \, d\tau, \quad 0 \leq t \leq T, \]

where \( B'(\cdot) \) is a standard white Gaussian noise process.

A. Signals and Signal Space

Suppose \( X(\cdot) \) is in the set \( \mathcal{L}^2[0, T] \) of finite-energy signals in the interval \([0, T]\). Let \( \{ \phi_m(t) \}_{m=1}^\infty \) be an orthonormal basis of \( \mathcal{L}^2[0, T] \). We may write

\[ X(t) = \sum_{m=1}^\infty X_m \phi_m(t), \quad W(t) = \sum_{m=1}^\infty W_m \phi_m(t) \]

where

\[ X_m = \int_0^T X(t) \phi_m(t)^* \, dt, \]

\( x^* \) is the complex conjugate of \( x \), and the \( \{ W_m \}_{m=1}^\infty \) are independent and identically distributed (iid), complex-valued, respectively.
circularly symmetric, Gaussian random variables with zero mean and unit variance.

The projection of the received signal onto the \(n\)-th basis function is
\[
Y_n = \int_0^T Y(t) \phi_n(t)^* \, dt
\]
(6)
\[
= \sum_{m=1}^\infty X_m \int_0^T \phi_m(t) \phi_n(t)^* \, dt + W_n
\]
(7)
\[
= \sum_{m=1}^\infty X_m \Phi_{mn} + W_n.
\]
(8)
The set of equations given by (3) for \(n = 1, 2, \ldots\) can be interpreted as the output of an infinite-dimensional multiple-input multiple-output channel, whose fading channel matrix is \(\Phi = [\Phi_{mn}]\).

**B. Receivers with Finite Time Resolution**

Consider a receiver whose time resolution is limited to \(\Delta\) seconds, in the sense that every projection must include at least a \(\Delta\)-second interval. More precisely, we set \(ML\Delta = T\), where \(M\) is the number of independent symbols transmitted in \([0,T]\) and \(L\) is the oversampling factor, i.e., the number of samples per symbol. The integrate-and-dump receiver with resolution time \(\Delta\) uses the basis functions
\[
\phi_m(t) = \begin{cases} 
1/\sqrt{\Delta}, & t \in [(m-1)\Delta, m\Delta) \\
0, & \text{elsewhere}
\end{cases}
\]
(9)
for \(m = 1, \ldots, ML\). With the choice (6), the fading channel matrix \(\Phi\) is diagonal and the channel’s output for \(n = 1, \ldots, ML\) is
\[
Y_n = X_n \frac{1}{\Delta} \int_{(n-1)\Delta}^{n\Delta} e^{j\Theta(t)} \, dt + W_n
\]
\[
= X_n e^{j\Theta((n-1)\Delta)} \frac{1}{\Delta} \int_{(n-1)\Delta}^{n\Delta} e^{j(\Theta(t) - \Theta((n-1)\Delta))} \, dt + W_n
\]
\[
\overset{D}{=} X_n e^{j\Theta_n} \frac{1}{\Delta} \int_0^{\Delta} e^{j\gamma \sqrt{\Delta} B_n(t/\Delta)} \, dt + W_n
\]
\[
\overset{(a)}{=} X_n e^{j\Theta_n} e^{j\gamma \sqrt{\Delta} B_n(t/\Delta)} \, dt + W_n
\]
\[
= X_n e^{j\Theta_n} F_n + W_n,
\]
where we have used the notation \(\Theta_n = \Theta((n-1)\Delta)\) and \(F_n = \int_0^1 e^{j\gamma \sqrt{\Delta} B_n(t)} \, dt\). In (10) we have used (3), the property \(B(t/T) - B((n-1)\Delta/T) \overset{D}{=} B(t/T - (n-1)\Delta/T)\), the substitution
\[
t = - (n-1)/\Delta
\]
\[
B_n(u/T) \overset{\Delta}{=} B(u/T - (n-1)\Delta/T),
\]
(12)
and the property \(\sqrt{T} B_n(t/\Delta) \overset{D}{=} \sqrt{\Delta} B_n(t/\Delta)\). Finally, in step (a) we have used the substitution \(t \overset{\Delta}{=} t/\Delta\).

Since the oversampling factor is \(L\), we have \(X_{kL+1} = X_{kL+2} = \ldots = X_{kL+L}\) for \(k = 0, \ldots, M-1\), and we can write the model (11) as
\[
Y_n = \sum_{m=1}^\infty X_m e^{j\Theta_n} F_n + W_n
\]
(13)
for \(n = 1, \ldots, ML\).

The vectors \(X_k^{ML}\), \(F_k^{ML}\) and \(W_k^{ML}\) are independent of each other. The variables \(\{X_k\}_{k=1}^M\) are chosen as iid with zero mean and variance \(E[|X_n|^2]\), and the average power constraint is
\[
E \left[ \frac{1}{T} \int_0^T |X(t)|^2 \, dt \right] = \frac{1}{ML\Delta} \sum_{n=1}^{ML} E[|X_n|^2]
\]
\[
= \frac{E[|X_n|^2]}{\Delta} \leq \mathcal{P}.
\]
(14)
Since we set the power spectral density of \(W\) to 1, the power \(\mathcal{P}\) is also the SNR, i.e., \(\mathcal{SNR} = \mathcal{P}\).

Using (3), the variables \(\Theta_k^{ML}\) follow a discrete-time Wiener process:
\[
\Theta_n = \Theta_{n-1} + N_{n-1}, \quad n = 1, \ldots, M,\]
(15)
where the \(N_n\)’s are iid Gaussian variables with zero mean and variance \(\gamma^2\Delta\). The fading variables \(F_n\)’s are complex-valued and iid, and \(F_n\) is independent of \(\Theta_n\). In other words, \(F_n\) is correlated only to \(N_n\), and is independent of the vector \((N_{n-1}, \ldots, N_{ML})\).

Note that for any finite \(\Delta\), or equivalently for any finite oversampling factor \(L\), the vector \(Y_k^{ML}\) does not represent a sufficient statistic for the detection of \(X\) given \(Y\) in the model (1).

**III. LOWER BOUND ON CAPACITY**

We compute a lower bound to the capacity of the continuous-time Wiener phase noise channel (3). For notational convenience, we use the following indexing for \(i = 1, \ldots, L\) and \(k = 1, \ldots, M;\)
\[
Y_{(k-1)L+i} = X_k e^{j\Theta((k-1)L+i)} F_{(k-1)L+i} + W_{(k-1)L+i},
\]
(16)
and we group the output samples associated with \(X_k\) in the vector \(\mathbf{Y}_k = [Y_{(k-1)L+1}, \ldots, Y_{(k-1)L+L}]\).

The capacity is defined as
\[
C(\mathcal{SNR}) = \lim_{M \rightarrow \infty} \frac{1}{M} \sup \sum_{i=1}^M I(X_k^M; Y_k^M)
\]
(17)
where the supremum is taken among the distributions of \(X_k^M\) such that the average power constraint (3) is satisfied.

The mutual information rate can be lower-bounded as follows:
\[
\frac{1}{M} \sum_{k=1}^M I(X_k^M; Y_k^M) = \frac{1}{M} \sum_{k=1}^M I(X_k; Y_k^M | X_k^{k-1})
\]
\[
\overset{(a)}{=} \frac{1}{M} \sum_{k=1}^M \sum_{i=1}^M I(X_k; Y_k^M | X_k^{i-1}) + I(X_k; Y_k^M | X_k^{i-1}, | X_k^i)
\]
\[
\overset{(b)}{=} \frac{1}{M} \sum_{k=1}^M \sum_{i=1}^M I(X_k^2; |Y_k^i|^2) + I(X_k; Y_k^M | X_k^{i-1}, | X_k^i)
\]
\[
= I(X_1^2; |Y_1|^2) + I(X_1; Y_0^1, X_0_1 | X_1)\]
(18)
where step (a) follows by polar decomposition of $X_k$, step (b) holds by a data processing inequality, by reversibility of the map $x \mapsto x^2$ for non-negative reals, and because $X_k^{k-1}$ is independent of $(X_k, Y_k)$. Finally, the last equality follows by stationarity of the processes.

### A. Amplitude Modulation

By choosing a specific input distribution that satisfies the average power constraint we always get a lower bound on the mutual information, so we choose the input distribution as

$$p_{|X_k|^2}(x) = \begin{cases} \frac{1}{\lambda} \exp \left( -\frac{x - \Delta - t}{\lambda} \right) & x \geq \Delta - t \\ 0 & \text{elsewhere} \end{cases}$$

(19)

where $\lambda = SNR \Delta - \Delta - t > 0$ with $t > 0$. Note that with this choice the average power constraint is satisfied with equality, i.e., $E[|X_k|^2] = SNR \Delta$.

Similar to the method used in [3], we give here a lower bound to the first term on the right hand side (RHS) of (18) in the form

$$I_1 \geq E[-\ln q_V(V)] - E[-\ln q_{V|X_k^2}(V|X_k^2)]$$

(20)

where $V = ||Y_1||^2$ and

$$q_{V}(v) = \int_0^\infty p_{|X_k|^2}(x)q_{V|X_k^2}(v|x)\,dx.$$  

(22)

Specifically, we choose the auxiliary channel distribution as

$$q_{V|X_k^2}(v|x) = \frac{1}{\sqrt{\pi \nu x}} \exp \left( -\frac{(v - L(1 + xE[G]))}{\nu x} \right)$$

(23)

where $G = ||F_1||^2/L$ and $\nu > 0$, for which we have

$$E[-\ln q_{V|X_k^2}(V|X_k^2)] = \frac{1}{2} \ln \left( \frac{\pi \nu x} {2} \right) + \frac{1}{2} E[\ln(|X_k|^2)] + \frac{L}{\nu} \left( \frac{E[|X_k|^2]}{\Delta} \ln [G] + 2E[G] + E \left[ \frac{1}{|X_k|^2} \right] \right)$$

$$\leq \frac{1}{2} \ln \left( \frac{\pi \nu x} {2} \right) + \frac{\Delta - t}{2\lambda} + \frac{L}{\nu} \left( SNR \cdot \ln [G] + 2 + \Delta' \right)$$

(23)

where the inequality is due to $E[G] \leq 1$, $E[|X_k|^2] \leq SNR \Delta$, the bound $E[|X_k|^2] \leq SNR \Delta$, and the support of $|X_k|^2$, and

$$E[\ln |X_k|^2] = \int_0^\infty \frac{1}{\lambda} \exp \left( -\frac{x - \Delta - t}{\lambda} \right) \ln(x)\,dx$$

$$= \ln \lambda + \int_0^\infty \frac{1}{\exp \left( -\frac{(u - \Delta - t)}{\lambda} \right) } \ln(u)\,du$$

$$\leq \ln \lambda + \frac{\Delta - t}{\lambda}.$$  

(24)

By substituting (22) and (19) into (23), and by following similar steps to those of [3], we get

$$E[-\ln q_{V}(V)] \geq -\frac{\Delta - t}{\lambda} + \frac{1}{2} \ln (L^2 \mu^2 \lambda^2 + \lambda \nu).$$

(25)

By putting together (23) and (25) we obtain

$$I_1 \geq -\frac{3\Delta - t}{2\lambda} - \frac{1}{2} \ln (L^2 \mu^2 \lambda^2 + \lambda \nu) + \frac{1}{2} \ln (\pi \nu x)$$

$$- \frac{L}{\nu} \left( SNR \cdot \ln |G| + 2 + \Delta' \right).$$  

(26)

In the limit of large time resolution we have

$$\lim_{\Delta \to 0} \frac{\ln |G|}{\Delta} = -\frac{\mu^2 \lambda^2}{4} + \frac{\nu}{4}.$$

(27)

Now we let the time resolution grow as a power of the SNR, i.e., $\Delta = [\text{SNR}^\alpha]$, and the parameter $\nu = \rho \Delta^{-2}$, with $\rho > 0$. By using (27) into (26), in order to find a tight bound in the interval $1/3 \leq \alpha \leq 1$ we need to satisfy the condition $\alpha < 1/(t+1)$ and $\beta \geq 1$. The tightest bound is obtained with $\beta = 1$ and $\rho = 4$:

$$\lim_{SNR \to \infty} \left\{ I_1 - \frac{1}{2} \ln (\text{SNR}) \right\} \geq -\frac{1}{2} \ln \left( \frac{2 \pi \gamma^2 e}{45} \right).$$

(28)

### B. Phase Modulation

The second term in the RHS of (18) can be lower-bounded as follows

$$I (\angle X_1; Y_0^1 | X_0, |X_1|) \geq I (\angle X_1; \Phi | X_0, |X_1|)$$

$$\geq E[-\ln q_{\Phi|X_0,|X_1}(\Phi|X_0, |X_1|)] - E[-\ln q_{\Phi|X_0^2}(\Phi|X_0^2)]$$

(30)

where step (a) is due to a data processing inequality with

$$\Phi = \angle (Y_k^1 Y_0^1 e^{-j \angle X_0^*})$$

$$\angle X_1 \oplus \angle (|X_1| F_1 + W_1) \oplus \angle (|X_0^2| F_0^2 e^{j X_0} + W_0^*),$$

(31)

and the last inequality follows by choosing the auxiliary channel

$$q_{\Phi|X_0^2}(\phi|x_0^2) = \frac{\exp(\phi \cos(\phi - \angle x_1))}{2\pi I_0(\zeta)}$$

(32)

where $I_0(\cdot)$ is the zero-th order modified Bessel function of the first kind, and $\zeta$ is a positive real number. Since we assume an uniform input phase distribution, the output distribution is also uniform:

$$q_{\Phi|X_0,|X_1}(\phi|x_0, |x_1|) = \int_0^{2\pi} q_{\Phi|X_0^2}(\phi|x_0^2) \frac{1}{2\pi} d\angle x_1 = \frac{1}{2\pi}.$$  

(33)

Details are provided in the extended version of the paper.
Using (32), the second term in the RHS of (30) can be upper-  
bounded as follows for any \( \Delta \leq \Delta < \infty \):

\[
E \left[ - \ln q_{\Phi|X_0} (\Phi|X_0) \right] = \ln (2\pi I_0(\zeta)) - \zeta E [\cos(\Phi - \angle X_1)] \\
\leq \ln (\pi \sqrt{\pi}) + \frac{1}{2} \ln \left( \frac{1}{\zeta} \right) + \zeta \rho \\
= \frac{1}{2} \ln \left( 2\pi^3 e^\rho \right)
\]

where the inequality is due to \( I_0(\zeta) \leq \sqrt{\pi/2}.e^\zeta/\sqrt{\zeta} \) derived in [11], Lemma 2, and from the result of Appendix A with

\[
\rho = 1 - E [F_0 e^{-jN_0}] E [F_1] + 2e^{-3j\gamma^2/4} E [|X_1|^{-2}] K_\Delta
\]

where \( K_\Delta > 1 \) is a finite number.\(^3\) The last step in (34) is obtained by choosing \( \zeta = (2\rho)^{-1} \).

In the limit of large time resolution we have

\[
\lim_{\Delta \to 0} \left\{ \frac{\rho}{\Delta} - 2K_\Delta \Delta^{1-1} \right\} \leq \frac{2}{3} \gamma^2
\]

where the inequality follows from the bound \( E [|X_1|^{-2}] \leq \Delta^t \). Choosing \( t = 1 \) and putting together (32) and (33)–(36) we get

\[
\lim_{\Delta \to 0} \left\{ I_L + \frac{1}{2} \ln (\Delta) \right\} \geq \frac{1}{2} \ln \left( \frac{3}{\pi e(\gamma^2 + 3K_\Delta)} \right),
\]

and letting the time resolution grow as a power of the SNR, i.e., \( \Delta^{-1} = [\text{SNR}^\alpha] \), for \( 0 < \alpha \leq 1/2 \) we have

\[
\lim_{\text{SNR} \to \infty} \left\{ I_L - \frac{\alpha}{2} \ln (\text{SNR}) \right\} \geq \frac{1}{2} \ln \left( \frac{3}{\pi e(\gamma^2 + 3K_\Delta)} \right).
\]

IV. DISCUSSION

As a byproduct of (28), (29), and (38), a lower bound to the capacity pre-log is

\[
\lim_{\text{SNR} \to \infty} \frac{C(\text{SNR})}{\ln (\text{SNR})} \geq \begin{cases} 
2\alpha & 0 < \alpha \leq 1/3 \\
(1 + \alpha)/2 & 1/3 \leq \alpha \leq 1/2 \\
3/4 & 1/2 \leq \alpha < 1.
\end{cases}
\]

Figure 1 shows the lower bounds on the capacity pre-log versus the parameter \( \alpha \), as reported by (39). The contributions of amplitude and phase modulation are also shown separately: Amplitude modulation reaches full degrees of freedom by sampling more than \( \sqrt{\text{SNR}} \) samples per symbol, while phase modulation achieves at least half of the available degrees of freedom by using a time resolution that scales as \( 1/\sqrt{\text{SNR}} \).

The input distribution that achieves the capacity lower bound is uniform in phase and the square amplitude is distributed as a shifted exponential (19). The statistic used for detecting \( |X_k| \) is \( |Y_k| \), and the one used for detecting \( \angle X_k \) is \( \angle \left( Y_{(k-1)L+1} Y_{(k-1)L} e^{-j\angle X_{k-1}} \right) \).

\(^3\)For example, choosing \( \gamma^2 \Delta = 0.01 \) gives \( K_\Delta = 8.1353 \). See the extended version of the paper for a detailed derivation.
$W_1$. A lower bound to (41) is given by

$$E \left[ \mathcal{R}\{ e^{jL(\langle X_1 F_1 + W_1 \rangle)} \} \right] \geq E \left[ \mathcal{R}\{ F_1 \} \cos(\langle X_1 F_1 + W_1 \rangle) \right] \tag{a}$$

$$\geq E \left[ \mathcal{R}\{ F_1 \} 1(\mathcal{R}\{ F_1 \} \geq 0) \left( 1 - \frac{1}{|X_1 F_1|^2} \right) \right] \tag{b}$$

$$+ E \left[ \mathcal{R}\{ F_1 \} 1(\mathcal{R}\{ F_1 \} < 0) \right] \geq E \left[ \mathcal{R}\{ F_1 \} - 1(\mathcal{R}\{ F_1 \} \geq 0) \right] \frac{1}{|X_1 F_1|^2} \tag{c}$$

$$\geq E \left[ \mathcal{R}\{ F_1 \} - \frac{1}{|X_1 F_1|^2} \right] \geq \frac{2}{\gamma^2 \Delta} (1 - e^{-\gamma^2 \Delta/2}) = E \left[ \frac{1}{|X_1|^2} \right] K_\Delta \tag{d}$$

where step (a) holds because $|F_1| \leq 1$, (b) follows by $\cos(x) \leq 1$ and by

$$E \left[ \cos(\langle \rho + W_1 \rangle) \right] \geq 1 - \frac{1}{\rho^2}, \quad \rho > 0, \quad \tag{43}$$

step (c) because $\mathcal{R}\{ F_1 \} \leq 1$, step (d) is obtained by subtracting $E \left[ 1(\mathcal{R}\{ F_1 \} < 0) |X_1 F_1|^2 \right]$, and the final inequality uses $E \left[ |F_1|^2 \right] \leq K_\Delta$, for a finite suitable $\Delta$.

Following an analogous derivation used for finding (42), for the second factor on the RHS of (40) we have

$$E \left[ R \{ e^{jL(\langle X_0 F_0 e^{-jN_0} + W_0 \rangle)} \} \right] \geq E \left[ R \{ F_0 e^{-jN_0} \} - \frac{1}{|X_0 F_0|^2} \right] \geq \sqrt{\frac{2\pi}{\gamma^2 \Delta}} \text{erf} \left( \sqrt{\frac{\gamma^2 \Delta}{8}} \right) e^{-3\gamma^2 \Delta/8} - E \left[ \frac{1}{|X_0|^2} \right] K_\Delta \tag{44}$$

where $\text{erf}(\cdot)$ is the error function, and the closed form for $E \left[ F_0 e^{-jN_0} \right]$ is provided in Appendix B. Using (42) and (43) into (40), with $R\{E\{F_1\}\} \leq 1, R\{E\{F_0 e^{-jN_0} \}\} \leq e^{-3\gamma^2 \Delta/8}$, $E\left[|X_0|^2\right] \geq 0$, and $E\left[|F_0|^2\right] \geq 0$, the final result is

$$E \left[ \cos(\Phi - \langle X_1 \rangle) \right] \geq E \left[ F_0 e^{-jN_0} \right] E \left[ F_1 \right] - 2e^{-3\gamma^2 \Delta/8} E \left[ \frac{1}{|X_1|^2} \right] K_\Delta \tag{45}$$

**APPENDIX B**

**EVALUATION OF $E \left[ F_0 e^{-jN_0} \right]$**

Knowing that $N_0 = \sigma \int_0^1 B(\tau) d\tau$ with $\sigma = \gamma \sqrt{\Delta}$, we compute

$$\text{Var} \left[ \sigma B(t) - N_0 \right] = \sigma^2 \text{Var} \left[ B(t) \right] + \text{Var} \left[ N_0 \right] - 2\sigma E \left[ B(t) N_0 \right] = \sigma^2 (t + 1) - 2\sigma^2 \int_0^1 E \left[ B(t) B(\tau) \right] d\tau = \sigma^2 (t^2 - t + 1) \tag{46}$$

where the last step follows from the property of Wiener processes $E \left[ B(t) B(\tau) \right] = \min\{t, \tau\}$. Thus we have

$$E \left[ F_0 e^{-jN_0} \right] = \int_0^1 E \left[ e^{j(\sigma B(t) - N_0)} \right] dt \tag{47}$$

where in step (a) we used the characteristic function of a Gaussian random variable, and in the last step we used (45).

**ACKNOWLEDGMENT**

L. Barletta and G. Kramer were supported by an Alexander von Humboldt Professorship endowed by the German Federal Ministry of Education and Research.

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Appendix C
Auxiliary channel

Choose the auxiliary channel distribution

$$q_{V|X_1} (v|x) = \frac{1}{\sqrt{2\pi \nu x}} \exp \left( -\frac{(v - L(1 + x\mu))^2}{\nu x} \right)$$  \hspace{1cm} (48)$$

where

$$V = \|Y_1\|^2 = |X_1|^2LG + Z_0 + 2|X_1|Z_1$$  \hspace{1cm} (49)$$

and

$$G = \frac{||F_i||^2}{L}, \quad Z_0 = ||W_1||^2, \quad Z_1 = \sum_{i=1}^{L} \Re \{ F_i W_i^* \}. \hspace{1cm} (50)$$

Squaring the output statistic gives

$$V^2 = |X_1|^4L^2G^2 + 4|X_1|^3LGZ_1 + 2|X_1|^2(2Z_1^2 + LGZ_0) + 4|X_1|Z_0Z_1 + Z_0^2,$$  \hspace{1cm} (51)$$

thus

$$(V - L(1 + |X_1|^2 \mu))^2$$

$$= V^2 - 2VL(1 + |X_1|^2 \mu) + L^2(1 + |X_1|^2 \mu)^2$$

$$= |X_1|^4L^2G^2 + 4|X_1|^3LGZ_1 + 2|X_1|^2(2Z_1^2 + LGZ_0) + 4|X_1|Z_0Z_1 + Z_0^2 + L^2(1 + |X_1|^2 \mu)^2$$

$$- 2(|X_1|^2LG + Z_0 + 2|X_1|Z_1)L(1 + |X_1|^2 \mu)$$

$$= |X_1|^4L^2(G - \mu)^2 + 2|X_1|^2(2Z_1^2 + L(Z_0 - L)(G - \mu)) + 4|X_1|Z_1(Z_0 - L) + (Z_0 - L)^2.$$  \hspace{1cm} (52)$$

Taking expectations gives

$$E \left[ \frac{(V - L(1 + |X_1|^2 \mu))^2}{\nu |X_1|^2} \right]$$

$$= \frac{L^2}{\nu} E \left[ |X_1|^2 \right] E \left[ (G - \mu)^2 \right] + \frac{4L}{\nu} E \left[ |X_1| \right] E \left[ Z_1(Z_0 - L) \right]$$

$$+ \frac{2}{\nu} \left( 2E \left[ Z_1^2 \right] + LE \left[ (Z_0 - L)(G - \mu) \right] \right)$$

$$+ \frac{4}{\nu} E \left[ \frac{1}{|X_1|} \right] E \left[ Z_1(Z_0 - L) \right] + \frac{1}{\nu} E \left[ \frac{1}{|X_1|^2} \right] E \left[ (Z_0 - L)^2 \right]$$

$$= \frac{L^2}{\nu} E \left[ |X_1|^2 \right] E \left[ (G - \mu)^2 \right] + \frac{2L}{\nu} E \left[ G \right] + \frac{L}{\nu} E \left[ \frac{1}{|X_1|^2} \right]$$  \hspace{1cm} (53)$$

which is minimized by choosing $\mu = E \left[ G \right]$.  

The conditional entropy is

$$E \left[ -\ln q_{V|X_1} (V | X_1^2) \right]$$

$$= \frac{1}{2} \ln (\pi \nu) + \frac{1}{2} \frac{E \left[ \ln (|X_1|^2) \right]}{\nu |X_1|^2} + \frac{1}{2} \left( \frac{E \left[ \frac{|X_1|^2}{\Delta} \right]}{\Delta} Var \left[ G \right] + 2E \left[ G \right] + E \left[ \frac{1}{|X_1|^2} \right] \right)$$

$$\leq \frac{1}{2} \ln (\pi \nu \lambda) + \frac{\Delta^t}{2\lambda}$$

$$+ \frac{L}{\nu} \left( SNR \cdot Var \left[ G \right] + 2E \left[ G \right] + E \left[ \frac{1}{|X_1|^2} \right] \right)$$  \hspace{1cm} (54)$$

where in the last inequality we have used the bound $E \left[ \ln \lambda \right]$.  

The output distribution is

$$q_{V}(v) = \int_{0}^{\infty} p_{X_1}(x) q_{V|X_1} (v|x) \, dx$$

$$= \int_{0}^{\infty} \frac{1}{\lambda} \exp \left( -\frac{x - 1}{\lambda} \right) \frac{1}{\sqrt{2\pi \nu x}} \exp \left( -\frac{(v - L(1 + x\mu))^2}{\nu x} \right) \, dx$$

$$\leq \exp \left( \frac{1}{\lambda} \right) \int_{0}^{\infty} \frac{1}{\lambda} \exp \left( -\frac{x}{\lambda} \right) \frac{1}{\sqrt{2\pi \nu x}} \exp \left( -\frac{(v - L(1 + x\mu))^2}{\nu x} \right) \, dx$$

$$= \exp \left( \frac{1}{\lambda} \right) \int_{0}^{\infty} \frac{1}{\lambda} \exp \left( -\frac{y}{\lambda} \right) \frac{1}{\sqrt{2\pi \nu y}} \exp \left( -\frac{(v - L - y)^2}{\nu y / (\lambda \nu)} \right) \, dy$$

$$= \exp \left( -\frac{1}{\lambda} \right) \frac{1}{\sqrt{2\pi \nu}} \exp \left( -\frac{v - L - |v - L|}{\sqrt{1 + \frac{\nu}{L^2\mu^2 \lambda}}} \right),$$  \hspace{1cm} (55)$$

and the entropy of the output of the auxiliary channel is

$$E \left[ -\ln q_{V|X_1} (V) \right]$$

$$= -\frac{\Delta^t}{\lambda} + \frac{1}{2} \ln (L\mu\lambda (L\mu + \nu / (L\mu)))$$

$$- \frac{2L}{\nu} \left[ E \left[ V - L \right] - E \left[ |V - L| \right] \right] \frac{1}{\sqrt{1 + \frac{\nu}{L^2\mu^2 \lambda}}}$$

$$\geq -\frac{\Delta^t}{\lambda} + \frac{1}{2} \ln (L\mu\lambda (L\mu + \nu / (L\mu)))$$

$$+ \frac{2L}{\nu} \left[ E \left[ V - L \right] \right] \frac{1}{\sqrt{1 + \frac{\nu}{L^2\mu^2 \lambda}}} - 1$$

$$\geq -\frac{\Delta^t}{\lambda} + \frac{1}{2} \ln (L^2\nu \mu^2 \lambda^2 + \lambda \nu).$$  \hspace{1cm} (56)$$

where step (a) holds because $E \left[ |\cdot| \right] \geq E \left[ \cdot \right]$, and the last inequality holds because of

$$E \left[ V - L \right] = E \left[ |X_1|^2 LG + Z_0 + 2|X_1|Z_1 - L \right]$$

$$= E \left[ |X_1|^2 \right] |F_1|^2 \geq 0.$$  \hspace{1cm} (57)$$
The lower bound to the mutual information rate for the amplitude modulation is

\[
I_{||} \geq \mathbb{E}[-\ln q_V(V)] - \mathbb{E}[-\ln q_V|X_1|^2(V|X_1|^2)] \\
\geq -\frac{3\Delta^{-t}}{2\lambda} + \frac{1}{2} \ln (L^2\mu^2\lambda^2 + \lambda
\nu) - \frac{1}{2} \ln (\pi
\lambda) \\
- \frac{L}{\nu} (\text{SNR} \cdot \text{Var}[G] + 2\mathbb{E}[G]) + \mathbb{E}\left[\frac{1}{|X_1|^2}\right]. \tag{58}
\]

In order to have a tight bound, we need to satisfy the constraints

\[
\begin{align}
1 - \alpha(t + 1) &> 0 \\
\alpha(1 - \beta) + 1 - 3\alpha &\leq 0 \\
\alpha(1 - \beta) &\leq 0 \\
\alpha(1 - \beta) - \alpha t &\leq 0 \\
1 + \alpha(1 - \beta) &> \alpha(t - \beta + 2)
\end{align}
\]

that reduce to

\[
\begin{align}
\alpha < 1/(t + 1) \\
\alpha(1 - \beta) + 1 - 3\alpha &\leq 0 \\
\alpha(1 - \beta) &\leq 0.
\end{align}
\]

Next we consider the two cases $1/3 < \alpha < 1$ and $0 < \alpha < 1/3$.

If $\alpha > 1/3$, i.e., $1 - 3\alpha < 0$, then we have to satisfy

\[
\begin{align}
\alpha < 1/(t + 1) \\
\beta &\geq 1.
\end{align}
\]

For large SNR we have

\[
\lim_{\Delta \to 0} \mathbb{E}[G] = 1, \quad \lim_{\Delta \to 0} \text{Var}[G] = \frac{\gamma^2}{45} \tag{60}
\]

that gives

\[
\lim_{\text{SNR} \to \infty} \frac{\mathbb{E}[|X_1|^2; V]}{\Delta^3} \geq \lim_{\text{SNR} \to \infty} \left\{ -\frac{3}{2} \frac{1}{\text{SNR}^{1-\alpha(t+1)} - 1} \\
\frac{1}{2} \ln \left(\frac{(\text{SNR} - \text{SNR}^{\alpha(t+1)})^2 + \rho(\text{SNR}^{1-\alpha(1-\beta)} - \text{SNR}^{\alpha(t+\beta)})}{\text{SNR}^{1+\alpha(\beta-1)} - 1}\right) \\
- \frac{1}{2} \ln (\pi \rho) \left(1 - \frac{1}{\text{SNR}^{1-\alpha(t+1)}}\right)\right\} \\
- \frac{1}{2} \text{SNR}^{\alpha(1-\beta)} \left(\text{SNR}^{1-3\alpha} - \frac{\gamma^2}{45} + 2 + \text{SNR}^{-\alpha t}\right) \tag{61}
\]

A LOWER BOUND TO $\mathbb{E}[\cos(\angle(\rho + W))]$

The pdf of $\Psi = \angle(\rho + W)$ is $\frac{1}{\sqrt{4\pi}} \rho^\alpha \cos(\psi)e^{-\rho^2\sin^2(\psi)} \text{erfc}(\rho \cos(\psi)) \tag{69}$

\[
p_\Psi(\psi) = \frac{1}{2\pi} e^{-\rho^2} + \frac{1}{\sqrt{4\pi}} \rho^\alpha \cos(\psi)e^{-\rho^2\sin^2(\psi)} \text{erfc}(\rho \cos(\psi))
\]
where \( \text{erf} : x \mapsto 1 - \text{erf}(x) \) is the complementary error function. A lower bound to \( E[\cos(\Psi)] \) is\(^6\)

\[
E[\cos(\Psi)] \geq \frac{1}{2} \int_0^\pi \cos(\psi) \rho g(\psi) \, d\psi
\]

\[
\geq \int_0^{\pi/2} \frac{\rho}{\sqrt{\pi}} \cos^2(\psi) e^{-\rho^2 \sin^2(\psi)} \text{erfc}(-\rho \cos(\psi)) \, d\psi
\]

\[
\geq \int_0^{\pi/2} \frac{\rho}{\sqrt{\pi}} \cos^2(\psi) e^{-\rho^2 \sin^2(\psi)} \text{erfc}(-\rho \cos(\psi)) \, d\psi
\]

\[
= \int_0^{\pi/2} \frac{2\rho}{\sqrt{\pi}} \cos^2(\psi) e^{-\rho^2 \sin^2(\psi)} \, d\psi - \frac{\sqrt{\pi}}{4} \rho e^{-\rho^2}
\]

(70)

where (a) follows by symmetry, (b) follows by using \( \int_0^\pi \cos(\psi) \, d\psi = 0 \), (c) holds because the integrand is non-negative over the interval \( [\pi/2, \pi] \), (d) holds because \( \text{erfc}(\rho \cos(\psi)) \geq 2 - e^{-\rho^2 \cos^2(\psi)} \) and \( \cos^2(\psi) e^{-\rho^2 \sin^2(\psi)} \geq 0 \), and finally the last step follows by direct integration.

We bound the integral in (70) as follows

\[
\int_0^{\pi/2} \frac{2\rho}{\sqrt{\pi}} \cos^2(\psi) e^{-\rho^2 \sin^2(\psi)} \, d\psi
\]

\[
\geq \frac{2\rho}{\sqrt{\pi}} \int_0^{\pi/2} (1 - \psi^2)e^{-\rho^2 \psi^2} \, d\psi
\]

\[
= \left( 1 - \frac{1}{2\rho^2} \right) \text{erf} \left( \frac{\pi \rho}{2} \right) + \frac{1}{2\sqrt{\pi} \rho} e^{-\pi^2 \rho^2/4}
\]

\[
\geq \text{erf} \left( \frac{\pi \rho}{2} \right) - \frac{1}{2\rho^2}
\]

(71)

where inequality (a) follows from \( \cos^2(\psi) e^{-\rho^2 \sin^2(\psi)} \geq 1 - \psi^2 e^{-\rho^2 \psi^2} \), and the last step holds because \( \text{erf}(\cdot) \leq 1 \) and the last term is non-negative. Substituting back into (70) yields

\[
E[\cos(\Psi)] \geq \text{erf} \left( \frac{\pi \rho}{2} \right) - \frac{1}{2\rho^2} + \frac{\sqrt{\pi}}{4} \rho e^{-\rho^2}
\]

\[
\geq 1 - e^{-\pi^2 \rho^2/4} - \frac{1}{2\rho^2} - \frac{\sqrt{\pi}}{4} \rho e^{-\rho^2}
\]

\[
\geq 1 - \frac{4}{\pi^2 \rho^2 e} - \frac{1}{2\rho^2} - \frac{\sqrt{\pi}}{4\rho^2} \left( \frac{3}{2e} \right)^{3/2}
\]

\[
\geq 1 - \frac{1}{\rho^2}
\]

(72)

where step (a) is due to \( \text{erf}(x) \leq e^{-x^2} \), and inequality (b) follows from \( \rho^3 e^{-\rho^2} \leq (3/(2e))^{3/2} \) and \( e^{-\pi^2 \rho^2/4} \leq 4/(\pi^2 \rho^2 e) \).

\(^6\)The proof is the one proposed in the Ph.D. thesis of H. Ghozlan.

**APPENDIX E**

**AN UPPER BOUND TO \( E[|F_1|^2] \)**

Denoting \( Z = |F_1| \), we compute an upper bound as follows

\[
E[Z^{-2}] = \lim_{\delta \to 0} \int_0^{\infty} \frac{1}{x^2} p_Z(x) \, dx
\]

\[
\leq \lim_{\delta \to 0} \left\{ -\Pr(Z \leq \delta) + \int_\delta^\infty \frac{2}{x^3} \Pr(Z \leq x) \, dx \right\}
\]

\[
\leq \int_0^e \frac{2}{x^3} \mathbb{E}[1(Z \leq x)] \, dx + \int_e^\infty \frac{2}{x^3} \mathbb{E}[1(Z \leq x)] \, dx
\]

\[
\leq \int_0^e \frac{2}{x^3} \mathbb{E}[g(Z)] \, dx + \frac{1}{e^2}
\]

(73)

where \( \varepsilon \) is a suitably chosen positive number, step (a) follows by integrating by parts, inequality (b) holds because the cumulative function is always positive, and the last inequality holds by choosing a function \( g(Z) \geq 1(Z \leq x) \) for the first integral (i.e., for \( 0 < x < \varepsilon \)) and the inequality \( \mathbb{E}[1(Z \leq x)] \leq 1 \) for the second integral, that can be computed in closed form.

As for the function \( g(Z) \) we choose

\[
g(Z) = a(1 - Z^2)^2 \left( 1 - Z^2 \rho_1 \right) \left( 1 - Z^2 \rho_2 \right)
\]

\[
= a \left[ 1 - Z^2 (1 + \rho_1 + \rho_2) + Z^4 (\rho_1 + \rho_2 + \rho_1 \rho_2) - Z^6 \rho_1 \rho_2 \right]
\]

(74)

a polynomial whose positive roots are \( \{1, \rho_1^{-1/2}, \rho_2^{-1/2}\} \), and with \( g(0) = a > 1 \). To guarantee the positivity of \( g(Z) \) for \( 0 \leq Z \leq 1 \), we can design the roots \( \rho_1^{-1/2} \) and \( \rho_2^{-1/2} \) to be greater than 1, hence \( \rho_1 \) and \( \rho_2 \) less than 1.

If the polynomial \( g(Z) \) satisfies the condition \( \mathbb{E}[g(Z)] = 0 \), then we have a finite bound in (73):

\[
E[Z^{-2}] \leq \frac{1}{e^2}
\]

(75)

Imposing the condition \( \mathbb{E}[g(Z)] = 0 \) in (74) gives

\[
\rho_2 = -\frac{1 + \mathbb{E}[Z^2]}{\mathbb{E}[Z^2] + \mathbb{E}[Z^4]} \left( 1 + \rho_1 - \mathbb{E}[Z^2 \rho_1] \right)
\]

(76)

We want \( \rho_2 \) to be positive, for this we distinguish two cases. In the first case we have both numerator and denominator of (76) positive, and this is satisfied if

\[
\rho_1 \geq \max \left\{ \frac{\mathbb{E}[1 - Z^2]}{\mathbb{E}[Z^2(1 - Z^2)]}, \frac{\mathbb{E}[Z^2(1 - Z^2)]}{\mathbb{E}[Z^4(1 - Z^2)]} \right\} \geq 1 \quad \text{(77)}
\]

so this situation is not wanted. The other case is where both numerator and denominator are negative, i.e., for

\[
\rho_1 \leq \min \left\{ \frac{\mathbb{E}[1 - Z^2]}{\mathbb{E}[Z^2(1 - Z^2)]}, \frac{\mathbb{E}[Z^2(1 - Z^2)]}{\mathbb{E}[Z^4(1 - Z^2)]} \right\}
\]

(78)

Moreover, we want \( \rho_2 \leq 1 \), and imposing this condition on (76) when numerator and denominator are negative means

\[
\rho_1 \geq \frac{\mathbb{E}[(1 - Z^2)^2]}{\mathbb{E}[Z^2(1 - Z^2)]}
\]

(79)
Conditions (78) and (79) can be numerically checked by considering that
\[ E[Z^2] = \frac{2}{\alpha^2} (-1 + e^{-\alpha t} + \alpha t) \]  
(80)
\[ E[Z^4] = \frac{1}{\alpha^4} \left( \frac{87}{2} - \frac{392}{9} e^{-\alpha} + \frac{1}{18} e^{-4\alpha} - 30\alpha + 8\alpha^2 \right) \]  
\[ - \frac{40}{3} e^{-\alpha} \]  
(81)
\[ E[Z^6] = \frac{1}{\alpha^6} \left( -100\alpha^3 e^{-\alpha} + 144\alpha^3 - \frac{3}{25} \alpha e^{-4\alpha} \right) \]  
\[ - \frac{11991}{8} e^{-\alpha} + 1499\alpha - \frac{1}{200} e^{-9\alpha} \]  
\[ - \frac{2123}{3} \alpha^2 e^{-\alpha} + \frac{4}{15} \alpha^2 e^{-4\alpha} - 792\alpha^2 \]  
(82)
where \( \alpha = \frac{\gamma^2 \Delta}{2} \). For example, for \( \alpha = 1.3 \) and \( \gamma \sqrt{\Delta} = 0.1 \) the roots are \( \rho_1^{-1/2} \gg 1, \rho_2^{-1/2} \approx 1.0001 > 1 \).

The last thing to check is that
\[ g(Z) \geq 1, \quad \text{for } 0 \leq Z \leq x \text{ and all } 0 \leq x \leq \epsilon. \]  
(83)
Studying the convexity of \( g(Z) \), we find that the function is concave in \( 0 \leq Z \leq \epsilon_\gamma \) with
\[ \epsilon_\gamma^2 = \frac{\rho_1 + \rho_2 + \rho_1 \rho_2}{5\rho_1 \rho_2} \sqrt{\left( \frac{\rho_1 + \rho_2 + \rho_1 \rho_2}{5\rho_1 \rho_2} \right)^2 - \frac{1 + \rho_1 + \rho_2}{15\rho_1 \rho_2}}. \]  
(84)
Moreover, we have \( g(\epsilon_1) = 1 \) with \( \epsilon_1^2 \) given by Cardano’s formula. Condition (84) is verified if \( \epsilon \leq \min\{\epsilon_1, \epsilon_\gamma\} \), because for all \( 0 \leq Z \leq \epsilon \) we can guarantee \( g(Z) \geq 1 \) thanks to \( g(0) \geq 1, g(\epsilon) \geq 1 \), and concavity for \( g(Z) \). For example, for \( \alpha = 1.3 \) and \( \gamma \sqrt{\Delta} = 0.1 \) we have \( \epsilon_1 \approx 0.3506 \) and \( \epsilon_\gamma \approx 0.5774 \), so we choose \( \epsilon = 0.3506 \). Using (75) this gives a bound \( E[|F|^{-2}] \leq 8.1353 \) for all \( \gamma \sqrt{\Delta} \leq 0.1 \).

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