Zheng, J., Rodriguez, S., Laurin, C., Baird, D., Trela-Larsen, L., Erzurumluoglu, M., Zheng, Y., White, J., Giambartolomei, C., Zabaneh, D., Morris, R., Kumari, M., Casas, J-P., Hingorani, A. D., Evans, D., Gaunt, T., & Day, I. (2017). HAPRAP: a haplotype-based iterative method for statistical fine mapping using GWAS summary statistics. *Bioinformatics*, 33(1), 79-86. https://doi.org/10.1093/bioinformatics/btw565

Publisher's PDF, also known as Version of record

License (if available): CC BY

Link to published version (if available): 10.1093/bioinformatics/btw565

Link to publication record in Explore Bristol Research

PDF-document

This is the final published version of the article (version of record). It first appeared online via Oxford University Press at http://dx.doi.org/10.1093/bioinformatics/btw565. Please refer to any applicable terms of use of the publisher.

University of Bristol - Explore Bristol Research

General rights

This document is made available in accordance with publisher policies. Please cite only the published version using the reference above. Full terms of use are available: http://www.bristol.ac.uk/red/research-policy/pure/user-guides/ebr-terms/
Genetics and population analysis

HAPRAP: a haplotype-based iterative method for statistical fine mapping using GWAS summary statistics

Jie Zheng1,2,*, Santiago Rodriguez1,2, Charles Laurin1,2, Denis Baird1, Lea Trela-Larsen2, Mesut A. Erzurumluoglu2,3, Yi Zheng4, Jon White5,†, Claudia Giambartolomei5,†, Delilah Zabaneh5,†, Richard Morris2, Meena Kumari5, Juan P. Casas5,6, Aroon D. Hingorani5,7, on behalf of the UCLEB Consortium, David M. Evans1,8, Tom R. Gaunt1,2,‡,* and Ian N. M. Day2,‡

1MRC Integrative Epidemiology Unit, School of Social and Community Medicine, Bristol BS8 6BN, UK, 2School of Social and Community Medicine, University of Bristol, Bristol BS8 2BN, UK, 3Department of Health Sciences, Genetic Epidemiology Group, University of Leicester, Leicester LE1 7RH, UK, 4Dedman College of Humanities and Sciences, Southern Methodist University, Dallas, TX 750235, USA, 5Department of Genetics, Environment and Evolution, University College London Genetics Institute, London WC1E 6BT, UK, 6Department of Primary Care & Population Health, University College London, Royal Free Campus, London NW3 2PF, UK, 7Centre for Clinical Pharmacology, University College London, London WC1E 6BT, UK, Division of Medicine, and 8University of Queensland Diamantina Institute, Translational Research Institute, Brisbane, Australia, QLD 4102

*To whom correspondence should be addressed.
†Equal contributorship in UCLEB. UCLEB members are listed in Supplementary Text S1.
‡The authors wish it to be known that, in their opinion, the last 2 authors should be regarded as Joint Last Authors.
Associate Editor: Gunnar Ratsch

Received on November 18, 2014; revised on April 29, 2016; accepted on August 26, 2016

Abstract

Motivation: Fine mapping is a widely used approach for identifying the causal variant(s) at disease-associated loci. Standard methods (e.g. multiple regression) require individual level genotypes. Recent fine mapping methods using summary-level data require the pairwise correlation coefficients ($r^2$) of the variants. However, haplotypes rather than pairwise $r^2$, are the true biological representation of linkage disequilibrium (LD) among multiple loci. In this article, we present an empirical iterative method, HAPlotype Regional Association analysis Program (HAPRAP), that enables fine mapping using summary statistics and haplotype information from an individual-level reference panel.

Results: Simulations with individual-level genotypes show that the results of HAPRAP and multiple regression are highly consistent. In simulation with summary-level data, we demonstrate that HAPRAP is less sensitive to poor LD estimates. In a parametric simulation using Genetic Investigation of ANthropometric Traits height data, HAPRAP performs well with a small training sample size ($N < 2000$) while other methods become suboptimal. Moreover, HAPRAP’s performance is not affected substantially by single nucleotide polymorphisms (SNPs) with low minor allele frequencies. We applied the method to existing quantitative trait and binary outcome meta-analyses (human height, QTc interval and gallbladder disease); all previous reported association signals were replicated and two additional variants were independently associated with human height. Due to the growing availability of summary level data, the value of HAPRAP is likely to

© The Author 2016. Published by Oxford University Press.

This is an Open Access article distributed under the terms of the Creative Commons Attribution License (http://creativecommons.org/licenses/by/4.0/), which permits unrestricted reuse, distribution, and reproduction in any medium, provided the original work is properly cited.

Bioinformatics, 33(1), 2017, 79–86
doi: 10.1093/bioinformatics/btw565
Advance Access Publication Date: 1 September 2016
Original Paper
increase markedly for future analyses (e.g. functional prediction and identification of instruments for Mendelian randomization).

**Availability and Implementation:** The HAPRAP package and documentation are available at http://apps.biocompute.org.uk/haprap/

**Contact:** jie.zheng@bristol.ac.uk or tom.gaunt@bristol.ac.uk

**Supplementary information:** Supplementary data are available at Bioinformatics online.

# 1 Introduction

Genome-wide association studies (GWAS) have identified thousands of single nucleotide polymorphisms (SNPs) associated with human complex traits and diseases (Hindorff et al., 2009; Manolio, 2010). To increase the power to detect small genetic effects associated with common complex traits, meta-analysis of multiple GWAS studies have also been conducted including blood lipids (Teslovich et al., 2010, Electrocardiographic (ECG) traits (Arking et al., 2006; Gaunt et al., 2012; Newton-Cheh et al., 2009; Pfeuffer et al., 2009) and human height (Wood et al., 2014) amongst others.

When a plausible hit has been identified within a GWAS, the challenge becomes one of determining the independent potentially causal SNP signals from a background of many correlated variants within the linkage disequilibrium (LD) block. A common strategy adopted is to take the top association signal to represent the association in a genomic region. However, this design does not take into account the possibility of multiple causal variants within a region, which will result in an underestimation of the total variation that could be explained at a locus (Yang et al., 2012). Statistical methods are available to identify independent hits; however these methods either require access to individual level data, or rely on pairwise LD estimates when summary statistics are used.

Conditional analysis is time consuming when individual level genotype data from several cohorts needs to be analysed separately and then combined in meta-analysis (Zheng et al., 2013). Providing the pairwise LD structure is consistent in samples from the same ethnic group (Ke et al., 2004), there are two approximate conditional analysis methods that can effectively use GWAS summary data: Genome-wide Complex Trait Analysis (GCTA) conditional and joint effect analysis (COJO) (Yang et al., 2012) and Sequential sentinel SNP Regional Association Plots (SSS-RAP) (Zheng et al., 2013).

COJO is a state-of-the-art method extending the scope of multiple regression to summary-level meta-analysis. COJO estimates the approximate joint SNP effects from summary statistics in a meta-analysis and LD information from an appropriate reference sample. SSSRAP is a numerical and graphical approach that transforms the marginal SNP effect of a sentinel SNP to the joint SNP effect of a test SNP through a $2 \times 2$ SNP-haplotypes matrix.

These existing approximate conditional analysis methods use pairwise correlation coefficients ($r^2$) between SNPs to represent LD structure in each associated region. However, when considering regions with three or more causal variants, utilizing allele frequencies and pairwise LD correlation may lose LD information. Three-locus systems which represent combinations of co-inherited alleles within the same chromosome, are a more biologically correct way to represent LD among multiple loci. Fine mapping using haplotypes will pick up the LD information that is not detected using pairwise LD measures.

To aid the ‘missing LD information’ problem, we propose an empirical iterative method HAPlotype Regional Association analysis Program (HAPRAP) to improve the accuracy of approximate conditional analysis using GWAS summary data. The important difference between HAPRAP and COJO is that the former estimates the joint SNP effects by using haplotypes (rather than pair-wise LD) estimated from a reference sample. We use both simulations and real-data from the British Women’s Heart Health Study (BWHHS) (Lawlor et al., 2003) to show that HAPRAP outperforms COJO on a range of performance measures. We applied the method to group-level QTc interval data from the UCL-LSHTM-Edinburgh-Bristol (UCLER) meta-analysis (Shah et al., 2013), with the haplotype information estimated from imputed genotype data from the BWHHS; and human height from the Genetic Investigation of ANthropometric Traits (GIANT) meta-analysis (Wood et al., 2014), with the haplotype information estimated from the Avon Longitudinal Study of Parents and Children (ALSPAC). Both cases suggest that HAPRAP has increased power for fine mapping compared to COJO. We extended HAPRAP to binary phenotypes and we illustrate this with an example of meta-analysis for gallbladder disease (GBD) SNP hits (Rodriguez et al., 2015).

# 2 Materials and methods

## 2.1 Overview of the methodology

We aim to combine summary level statistics with the full information from haplotypes (rather than using the traditional pairwise LD approach) to fine map genetic regions. Our algorithm iteratively updates haplotype effects based on haplotype frequencies and observed marginal SNP effects from meta-analyses to estimate the approximate joint SNP effect. This approach allows researchers to conduct conditional analysis more accurately without access to individual level genotypes.

### 2.1.1 Theory

The haplotype-based approach we propose in this article is closely related to a single regression model. In a single regression model, we treat the major allele as the baseline allele; and the minor allele as the effect allele. The marginal SNP effect refers to the effect estimate from an outcome $Y$ regressed on a single SNP (i.e. the allelic effect from a simple linear regression model). The joint SNP effect, which we aim to estimate, refers to the SNP effect obtained from $Y$ regressed on multiple SNPs within the region. The joint SNP effect is adjusted for the correlation with surrounding SNPs, whereas the marginal SNP effect is not.

A simple extension of the single regression model to multi-locus data is to integrate two popular haplotype-based analysis strategies together: (i) dichotomize haplotypes into two groups (Lin and Zeng, 2006) and (ii) treat each group as a bivariate allele (Purcell et al., 2007a).

Assume we obtain a SNP by haplotype matrix $M$, with $m_{ij} = 0$ or $1$, from a sample population, we split existing haplotypes into two groups to estimate the joint effect of SNP $j$:

\[
\begin{align*}
HE_j = \{ i : m_{ij} = 1 \} \\
HB_j = \{ o : m_{oj} = 0 \}
\end{align*}
\]
HAPRAP: a haplotype-based iterative method

Fig. 1. The SNP by haplotypes matrix for HAPRAP. The iteration of HAPRAP is built based on a matrix summarizing the haplotypes and haplotype frequencies for a certain population. ‘0’ in the matrix means the haplotype contains the baseline allele for the relevant SNP, whereas ‘1’ means the haplotype contains the effect allele for the relevant SNP. The small arrow (from left to right) is the marginal SNP effects estimation step. The large arrow (from right to left) is the haplotype effects adjustment step.

\[ \begin{align*}
\text{HE}_e &= \text{the set of haplotypes containing the effect allele of SNP } j; \\
\text{HB}_o &= \text{the set of haplotypes containing the baseline allele of SNP } j. 
\end{align*} \]

For example for SNP1 in Figure 1, HE_1 is the set of haplotypes containing the effect allele of SNP 1, whereas HB_1 is the set of haplotypes from Haplotypes 5–8, whereas HB_1 is the set of haplotypes from Haplotypes 1–4. We also split the haplotype frequencies into two groups based on the relevant haplotypes F_i and F_o.

We then define the estimated marginal SNP effect of a SNP j, U_j as:

\[ U_j = \frac{1}{f_{o,j}} - \frac{1}{f_{o,j}} \quad \text{if } j \in \text{HE}_e \text{ and } o \in \text{HB}_o, \quad (1) \]

where \( f_{o,j} \) is the average of the additive effect over the set of haplotypes HE_e (or HB_o). These additive haplotype effects can be transferred to joint SNP effects using a generalized inverse matrix approach. This extension is applicable to both linear and logistic regression models.

2.1.2 HAPRAP algorithm for estimating the joint SNP effect

As individual-level genotype data is usually not publicly available for GWAS meta-analysis, we cannot estimate haplotype effects by conducting a haplotype-based association analysis. Thus, we use an iterative method to estimate the haplotype effects from marginal SNP effects. The iteration involves four steps (Fig. 2):

Step 1: Setting initial values for joint SNP effects and haplotype effects transformation.
Step 2: The marginal SNP effects estimation.
Step 3: The haplotype effects adjustment.
Step 4: Convergence and the generalized inverse matrix approach.

Table 1 provides details of the notation used in describing our method.

Fig. 2. Schematic diagram of HAPRAP

Assuming an additive linear model, the initial estimated haplotype effect \( Z^{(0)} \) is the matrix product of \( M \) and \( V^{(0)} \) (Fig. 1):

\[ MV^{(0)} = Z^{(0)}. \]

Step 2. Marginal SNP effects estimation:

As mentioned in Equation (1), we define the marginal SNP effect as the difference between the sums of the additive effects of the two sets of haplotypes HE_e and HB_o.

Thus, for the \( g \) iteration, \( U_g \), estimated by counting the difference between the two groups of haplotype effects, \( Z_{g}^{(0)} \) and \( Z_{g}^{(0)} \), and standardized by the relevant haplotype frequencies, \( F_i \) and \( F_o \):

\[ U_g^{(0)} = \frac{1}{\sum_{i \in \text{HE}} F_i} \left( \sum_{i \in \text{HE}} F_i Z_{g}^{(0)} \right) - \frac{1}{\sum_{o \in \text{HB}} F_o} \left( \sum_{o \in \text{HB}} F_o Z_{g}^{(0)} \right). \]

We tested the reliability of Equation (3) by a simulation and found that given any set of joint SNP effects, application of Equation (3) never generated non-zero effect estimates for SNPs that were simulated to have truly null effects (Supplementary Text S2).
Step 3. Haplotype effects adjustment: the adjusted marginal SNP effects for iteration g, \(U^{(g)}\) are compared to the observed marginal SNP effects, O. Reconciling the difference between \(U^{(g)}\) and O is important because it equates the marginal SNP effects observed from the meta-analytic data with those that would arise under the distribution of haplotypes in the reference panel. The SNP with the greatest deviation, denoted \(x^{(g)}\), is adjusted for the next iteration \(g+1\), the other SNP effects remain the same:

\[
U^{(g+1)} = \begin{cases} 
U^{(g)} - O, & \text{where } j = x^{(g)} \\
U^{(g)} & \text{where } j \neq x^{(g)}
\end{cases}
\]  

(4)

Then the haplotype effect \(Z^{(g+1)}\) will be adjusted based on the change of \(U^{(g+1)}\). For haplotype \(k\), we get:

\[
Z_k^{(g+1)} = Z_k^{(g)} + U_j^{(g)}m_{kj} \text{ where } j = x^{(g)}.
\]  

(5)

Step 4. Convergence and the generalized inverse matrix approach: After the estimated marginal SNP effects, \(U^{(g)}\) converge to within 10 decimal places of the observed SNP effects, O, we stop the iteration. The joint SNP effects, \(V^{(g)}\), is estimated using the generalized inverse matrix approach:

\[
M^{-1}Z^{(g)} = V^{(g)}.
\]  

(6)

2.2.2 Stepwise backwards elimination

For each iteration, we estimate the standard errors (SE) of the estimated joint SNP effects, \(V^{(g)}\), using the parametric bootstrap approach. The joint SNP effects, \(V^{(g)}\), converge to within 10 decimal places of the observed SNP effects, O, we stop the iteration. The joint SNP effects, \(V^{(g)}\), is estimated using the generalized inverse matrix approach:

\[
M^{-1}Z^{(g)} = V^{(g)}.
\]  

(6)

2.2.3 HAPRAP availability

The HAPRAP software and a web-based instruction manual (developed using HTML and Cascading Style Sheets (CSS)) are available at http://apps.biocompute.org.uk/haprap.

2.3 Sample datasets

The real cases and simulated datasets we used for this analysis are explained in Supplementary Text S3.

2.4 Simulation framework and empirical comparison

Firstly, we simulated a pool of 100 000 individuals (details in Supplementary Text S3) and performed a series of simulations to test the influence of LD structure and sample size of reference panel. For each model explained in Supplementary Text S3 and Supplementary Table S1, we applied HAPRAP and COJO to the summary statistics and the genotypes of a specific reference panel. We also applied multiple regression using individual-level phenotypes and genotypes from the reference panel. For each method, the mean and SD of the joint SNP effect were estimated 1000 times. In addition, multiple regressions on the 100 000 individuals were conducted (Supplementary Text S3) and the resulting joint SNP effects were set as the gold standards. Mean square error (MSE) of the gold standard effect was used to measure the accuracy of each method.

Secondly, we performed a parametric simulation to test the influence of the sample size of a meta-analysis. The GIANT height meta-analysis data were used as the basis of this simulation (Wood et al., 2014). We selected 20 nearest SNPs from the ACAN region. ALSPAC pre-phased haplotypes of 8263 unrelated children were used to build a genotype pool for 253 288 individuals. We randomly selected 100 000, 50 000, 10 000, 3000, 2500, 1750 and 1000 individuals from the pool, comparing the performance of HAPRAP and COJO using multiple regression as the gold standard. 1000 replications were processed to estimate the MSE and SD of the MSE.

Thirdly, as an empirical comparison between HAPRAP and COJO, we explored these methods using real data from the BWHHS and the 1000 Genomes project (1000 Genome Project Consortium, 2010). Details of the performance comparisons are explained in Supplementary Text S4.

2.5 Case study for quantitative traits: GIANT height

We firstly applied HAPRAP to two meta-analyses. Details of these two case studies are explained in Supplementary Text S5. We further applied HAPRAP to summary-level data from the GIANT height meta-analysis (sample size 253 288). The pre-phased haplotypes of 8263 unrelated children from ALSPAC were used as the reference panel. Three genomic regions with more than one robust independent association signal were selected (Wood et al., 2014). All SNPs within these regions were selected (782 SNPs for ACAN, 1477 SNPs for ADAMTS17 and 1936 SNPs for PITCH1).

3 Results

3.1 Simulation and empirical comparison

Firstly, we fixed the sample size of the meta-analysis (\(N = 100\,000\)) and compared the performance of HAPRAP and COJO across different LD structures and different sample sizes of reference panel using a simulation data set (details in Supplementary Text S3). As shown in Supplementary Table S2, HAPRAP outperformed COJO under a variety of LD structures and was less sensitive to poor LD estimation.

In the two-SNP models with one causal SNP and one non-effect SNP, HAPRAP was slightly (up to 29%) more accurate than COJO.
across 16 models (Supplementary Fig. S2A and Supplementary Table S2A). Both methods performed well when the sample size of the reference panel was larger than 5000. When the sample size of the reference panel was limited to 500–1000, HAPRAP started to outperform COJO. On the other hand, considering the influence of LD structure, HAPRAP was up to 54% more accurate than COJO when LD between the two SNPs was extremely high ($r^2$ = 0.9).

In the three-SNP models with two causal SNPs and one non-effect SNP (Fig. 3 and Supplementary Fig. S2B), both methods performed relatively well when the sample size of the reference panel was larger than 5000 (although with more errors compared to the two-SNP models). However, both methods struggled to eliminate the non-effect SNP when the sample size of reference panel is <1000 and LD was very high amongst three SNPs. However, in a more realistic LD range ($r^2$ between each pair of SNPs from 0.1 to 0.5) and with a small reference sample size (N = 1000), HAPRAP was, on average, 63% more accurate than COJO (Supplementary Table S2B).

We demonstrated in this simulation that, when individual-level data is extremely limited, HAPRAP (using summary level data and a reference panel with a small number of individuals) is a better option than applying multiple regression to the reference panel with limited sample size (Supplementary Fig. S2C and Supplementary Table S2).

Secondly, in the parametric simulation using GIANT height data, we assumed perfect LD estimation and only consider the influence of sample size of the meta-analysis. As shown in Figure 4 and Supplementary Table S3, HAPRAP and COJO were close to optimal (Supplementary Text S6 explains the reason COJO is not perfectly optimal in this situation) when the sample size of the meta-analysis was large (N ≥ 10 000). When the training sample size was between 1750 and 5000, HAPRAP’s MSE was still under 0.1 while COJO became suboptimal.

Thirdly, we utilized individual-level data of ~2000 BWHHS individuals on a total of 115 SNPs to compare the accuracy of HAPRAP [haplotypes phased by both segmented haplotype estimation and imputation tool (SHAPEIT) (Delaneau et al., 2012) and PLINK] and COJO using multiple regression as the gold standard (Supplementary Table S4). The details of the comparison can be found in Supplementary Text S4. In summary, the comparisons suggested that HAPRAP was comparable to multiple regression when the individual-level genotypes are available for the entire cohort. In addition, HAPRAP was on average 10.86% more accurate than COJO when the sample size of the reference panel was extremely limited (Sample size <200).

3.2 Case study: GIANT meta-analysis of height

We further analysed three genomic regions reported to be associated with human height by the GIANT consortium. The original fine mapping analyses were processed using COJO, resulting in 18 associated SNPs with P-value ≤ $5 \times 10^{-8}$ at these 3 loci (Wood et al., 2014). Here, we applied HAPRAP to a total of 4195 SNPs using 8263 unrelated ALSPAC children as a reference panel. The allele frequencies of GIANT and the ALSPAC children were quite similar (Supplementary Table S5). As shown in Table 2, HAPRAP replicated all 18 previously reported association signals at these 3 loci (Table 2). Moreover, HAPRAP identified two novel signals, rs1529889 (an intronic variant in ADAMST17 with joint effect of 0.019) and rs357564 (a missense variant in PTCH1 with joint effect of −0.034), independently associated with height, (Table 2). As shown in Supplementary Table S6, these two SNPs are in low LD with independent SNPs in the same genomic region.

Surprisingly, when we applied COJO to the same data using a different reference panel (ALSPAC instead of ARIC), only 16 SNPs were significantly associated with height, leaving two SNPs unselected (Supplementary Table S5).

We also conducted two case studies of GBD and QTc intervals. Details of these case studies are in Supplementary Text S5.

4 Discussion

Meta-analysis summary association statistics are becoming more and more widely available to the scientific community (Bulik-Sullivan et al. 2014).
In this article, we introduced a novel approach for statistical fine mapping using meta-analysis summary statistics. The proposed method (HAPRAP) uses haplotypes to represent LD structure among multiple variants in a region. Using haplotypes has four significant advantages compared to existing conditional analysis methods that utilize pairwise correlation coefficients ($r^2$) between SNPs [such as COJO (Yang et al., 2012) and SSSRAP (Zheng et al., 2013)]:

1. It considers all loci simultaneously, rather than pairwise, thus it is less susceptible to poor LD estimates that occur if the reference LD structure does not closely match the populations studied in the GWAS data.
2. It is more accurate than COJO when the sample size of the meta-analysis is limited (e.g. $N \leq 5000$).
3. It is more accurate and powerful for regions with three or more independent signals. Compared to Bayesian fine mapping methods such as Probabilistic Annotation INtegraTOR (PAINTOR) (Kichaev et al., 2014; Kichaev and Pasaniuc 2015), Causal Variants Identification in Associated Regions (CAVIAR) (Hormozdiari et al., 2014) and CAVIAR Bayes factor (CAVIARBF) (Chen et al., 2015), HAPRAP does not require the user to specify the number of causal variants. This can impair the performance of CAVIARBF for cases where there are multiple causal variants (Kichaev et al., 2014). We observed a power improvement in our case study of human height (e.g. with 3+ independent signals within each associated region).
4. It is more accurate when analysing rare variants (i.e. minor allele frequency (MAF) $< 0.01$) than other methods using pairwise LD.

Our empirical demonstration using the 1000 Genomes Project (1000 Genome Project Consortium, 2010) data comparison is meaningful in three aspects: Firstly, high quality haplotypes data, which is used by HAPRAP, are now widely available and should have already been pre-phased within large-scale consortiums/cohorts such as the aforementioned 1000 Genomes Project and ALSPAC. Secondly, for researchers without individual-level genotype data, our method can give researchers a general profile of the potentially multiple associated SNPs in the region(s) of interest using the public available 1000 Genome Project data, although the errors of using the 1000 Genomes Project data as a reference panel were relatively large since the sample size is currently small. As more open access phased haplotype data becomes available with the publication of projects, such as UK10K (UK10K consortium, 2015), HAPRAP's accuracy advantage against COJO will increase. Thirdly, HAPRAP's performance advantage will be more apparent for GWAS studies with relatively smaller sample sizes, such as association analyses of DNA methylation with expensive or high-dimensional phenotypes [e.g. gene expression and methylation data (Gaunt et al., 2015; Shi et al., 2014)].

In the case study using summary statistics of GIANT data (Wood et al., 2014), we identified two additional variants, rs1529889 and rs357564, independently associated with human height. These findings could have been caused by the greater sample size of the reference panel using ALSPAC (8263) compared to ARIC (6654), rs357564 is a missense variant within PTCH1 and rs1529889 is an intronic variant within ADAMST17. rs357564 is predicted to be ‘functional’ by the prediction tool FATHMM (Shihab et al., 2015) and was reported to be associated with oral clefts, basal cell carcinoma and ameloblastoma (Begnini et al., 2010; Carter et al., 2010; Farias et al., 2012).

Rare variants are on average younger than common variants (Mathieson and McVean, 2014) and are more likely to be represented by longer haplotypes. Since HAPRAP uses haplotypes and COJO uses pairwise LD, we show HAPRAP may have a theoretical advantage over COJO in rare variant analyses. We performed a simulation for two SNPs with MAFs near 0.08 (Supplementary Table S7) and HAPRAP’s accuracy was higher than COJO in all conditions. Moreover, we highlighted a rare variant in *Apolipoprotein B* (*APOB*, rs41288783, as a proof-of-concept using real data (Supplementary Table S8). This SNP had a MAF of 0.0018 in BWHHS individuals. The HAPRAP estimate ($\beta = 0.705$) is very close to the gold standard results ($\beta = 0.731$), whereas the COJO estimate is considerably different from the gold standard ($\beta = 0.449$).

We recommend using pre-phased haplotypes as HAPRAP input. For a cohort without haplotype data, we recommend users phase haplotypes using tools such as SHAPEIT (Delaneau et al., 2012), BEAGLE (Browning and Browning, 2009), IMPUTE2 (Howie et al., 2009) and Markov Chain Haplotypeing algorithm (MACH) (Li et al., 2010) rather than PLINK (Purcell et al., 2007b). PLINK haplotype phasing function uses an E-M algorithm, which is only accurate and fast when a small number of SNPs ($N < 10$) are included (Browning and Browning, 2011).

We also suggest controlling for collinearity before utilizing HAPRAP. If SNPs with high variance inflation factor (VIF) values are included, HAPRAP (and other tools) will return extremely
large betas for a pair of SNPs. Practically, it is necessary to remove SNPs with VIF higher than seven before applying HAPRAP.

HAPRAP requires more time than COJO to finalize the step-wise elimination process. There are several reasons: firstly, phasing haplotypes is time consuming; secondly, it is time consuming to determine the SE of the joint SNP effects using our bootstrap method (simHAPRAP). However, the whole process does not usually take more than an hour.

HAPRAP was originally designed for regional fine mapping, so it is more suitable for moderately small numbers of markers and computationally very fast when the number of SNPs in each test is 10 or fewer. To fit the HAPRAP framework to fine map the whole genome, we recommend splitting regions with large numbers of SNPs into smaller chunks (up to 20 SNPs in each chunk) before running HAPRAP. In the GIANT height example, we split the genomic regions based on recombination hotspots, since LD patterns are directly related to the underlying recombination process, which is a more reasonable option compared to the physical distance used by COJO. This can help reduce the run time of HAPRAP substantially.

Algorithms are often used effectively where the biological model is well understood, but the statistical model is too complex to generalize to all scenarios. For instance, a recent fine mapping method, probability identification of causal SNPs (PICS), used an empirical constant in its core algorithm to estimate the expected mean of the association signal at a SNP (Farh et al., 2015). HAPRAP interprets a complex biological concept, haplotype effects, using a simple idea stemming from allelic association analyses and extending it to the haplotype model. The side effect is that an asymptotic analysis of convergence may not be possible, thus we cannot exclude the possibility that HAPRAP will not converge in some situations. However, in the hundreds of thousands of simulations and real case examples we have tested, we did not find any situation where HAPRAP did not converge.

In a recent review paper (Spain and Barrett, 2015), fine mapping methods were classified into two groups: (i) methods for triaging variants based on P-values or LD with the lead SNP, which includes classic conditional analysis and approximate methods such as COJO and HAPRAP; (ii) Bayesian methods that assign posterior probabilities of membership in causal models to each SNP, such as PAINTER, CAVIAR, CAVIARBF and the most recent software, FINEMAP (Benner et al., 2016). Compared to CAVIARBF, FINEMAP used a new search algorithm and so is much faster and overcomes the limitation of situations where there are more than three causal variants in a genomic region. In addition, for the above Bayesian methods (with the exception of FINEMAP), a parameter must be set for the number of causal SNPs (Spain and Barrett, 2015). It has been shown that specifying this value to one can impair performance in cases where there are two or more causal variants (Kichaev et al., 2014). Based on this we consider HAPRAP and these Bayesian methods as complementary. It would be interesting to explore the potential of integrating the HAPRAP methods with these Bayesian algorithms to develop more powerful fine mapping methods in the future.

In conclusion, with increasing numbers of publicly available meta-analysis summary statistics, the value of HAPRAP is likely to be demonstrated in four ways: (i) for fine mapping both common and rare variants and identifying additional variants independently associated with complex traits; (ii) it can be used as a variable selection method for two-sample Mendelian randomization; (iii) to build genome-wide allelic scores of biological intermediates for mapping the phenotype (Evans et al., 2013); (iv) to provide a solid platform for the functional annotation of causal variants using prediction tools such as FATHMM (Supplementary Text S7).

Acknowledgement

We thank Frank Dudbridge for helpful advice on the article.

Funding

This work was financially supported by MRC Integrative Epidemiology Unit at the University of Bristol (MC_UU_12013/1-9). The funding information is listed in Supplementary Text S1.

Conflict of Interest: none declared.

References

1000 Genomes Project Consortium. et al. (2010) A map of human genome variation from population-scale sequencing. Nature, 467, 1061–1073.
Arking,D. E. et al. (2006) A common genetic variant in the NOS1 regulator NOS1AP modulates cardiac repolarization. Nat Genet, 38, 644–651.
Beghini,A. et al. (2010) PTCH1 gene haplotype association with basal cell carcinoma after transplantation. Br J Dermatol., 163, 364–370.
Benner,C. et al. (2016) FINEMAP; efficient variable selection using summary data from genome-wide association studies. Bioinformatics, 32, 1493–501.
Browning,B.L., and Browning,S.R. (2009) A unified approach to genotype imputation and haplotype-phase inference for large data sets of trios and unrelated individuals. Am. J. Hum Genet., 84, 210–223.
Browning,SR., and Browning,B.L. (2011) Haplotype phasing: existing methods and new developments. Nat Rev Genet, 12, 703–714.
Bulik-Sullivan,B. et al. (2015a) An atlas of genetic correlations across human diseases and traits. Nat Genet., 47, 1236–1241.
Bulik-Sullivan,B.K. et al. (2015b) LD Score regression distinguishes confounding from polygenicity in genome-wide association studies. Nat Genet., 47, 291–295.
Carter,T.C. et al. (2010) Testing reported associations of genetic risk factors for oral clefts in a large Irish study population. Birth Defects Res A Clin Mol Teratol, 88, 84–93.
Chen,W. et al. (2015) Fine mapping causal variants with an approximate Bayesian method using marginal test statistics. Genetics, 200, 719–736.
Delaneau,O. et al. (2012) A linear complexity phasing method for thousands of genomes. Nat Methods, 9, 179–181.
Evans,D.M. et al. (2013) Mining the human phenotype using allelic scores that index biological intermediates. PLoS Genet, 9, e1003919.
Farias,L.C. et al. (2012) Loss of heterozygosity of the PTCH gene in ameloblastoma. Hum Pathol., 43, 1229–1233.
Farh,K.K. et al. (2015) Genetic and epigenetic fine mapping of causal autoimmune disease variants. Nature, 518, 337–343.
Finucane,H.K. et al. (2013) Partitioning heritability by functional annotation using genome-wide association summary statistics. Nat Genet., 47, 1228–1235.
Gaunt,T.R. et al. (2012) Integration of genetics into a systems model of electrophysiological traits using HumanCVD BeadChip. Circ Cardiovasc Genet., 5, 630–638.
Gaunt,T.R. et al. (2013). Systematic identification of methylation quantitative trait loci across the human life course. Genome Biol., 17, 61.
Hindorff,L.A. et al. (2009) Potential etiologic and functional implications of genome-wide association loci for human diseases and traits. Proc Natl Acad Sci USA, 106, 9362–9367.
Hormozdiani,F. et al. (2014) Identifying causal variants at loci with multiple signals of association. Genetics, 198, 497–508.
Howie,B.N. (2009) A flexible and accurate genotype imputation method for the next generation of genome-wide association studies. PLoS Genet, 5, e1000529.
Ke,N. et al. (2004) Efficiency and consistency of haplotype tagging of dense SNP maps in multiple samples. Hum Mol Genet., 13, 2557–2565.
Kichaev,G. et al. (2014) Integrating functional data to prioritize causal variants in statistical fine mapping studies. PLoS Genet, 10, e1004722.
Kichaev,G., and Pasaniuc,B. (2015) Leveraging functional-annotation data in trans-ethnic fine-mapping studies. Am. J. Hum Genet., 97, 260–271.
Lawlor, D.A. (2003) Association between leg length and offspring birthweight: partial explanation for the trans-generational association between birthweight and cardiovascular disease: findings from the British Women’s Heart and Health Study. *Paediatr Perinat Epidemiol*, 17, 148–155.

Li, Y. et al. (2010) MaCH: using sequence and genotype data to estimate haplotypes and unobserved genotypes. *Genet. Epidemiol.*, 34, 816–834.

Lin, D.Y., and Zeng, D. (2006) Likelihood-based inference on haplotype effects in genetic association studies. *J. Am. Stat. Assoc.*, 101, 89–104.

Manolio, T.A. (2010) Genomewide association studies and assessment of the risk of disease. *N. Engl. J. Med.*, 363, 166–176.

Mathieson, I., and McVean, G. (2014). Demography and the age of rare variants. *PLoS Genet.*, 10, e1004528.

Newton-Cheh, C. et al. (2009) Common variants at ten loci influence QT interval duration in the QTGEN Study. *Nat. Genet.*, 41, 399–406.

Pasaniuc, B. et al. (2014) Fast and accurate imputation of summary statistics enhances evidence of functional enrichment. *Bioinformatics*, 30, 2906–2914.

Pfeufer, A. (2009) Common variants at ten loci modulate the QT interval duration in the QTSCD Study. *Nat. Genet.*, 41, 407–414.

Pierce, B.L., and Burgess, S. (2013) Efficient design for Mendelian randomization studies: subsample and 2-sample instrumental variable estimators. *Am. J. Epidemiol.*, 178, 1177–1184.

Purcell, S. et al. (2007a) WHAP: haplotype-based association analysis. *Bioinformatics*, 23, 255–256.

Purcell, S. et al. (2007b) PLINK: a tool set for whole-genome association and population-based linkage analyses. *Am. J. Hum. Genet.*, 81, 559–575.

Robinson, W.P. et al. (1991) Three-locus systems impose additional constraints on pairwise disequilibria. *Genetics*, 129, 925–930.

Rodriguez, S. et al. (2015) Lipids, obesity and gallbladder disease in women: insights from genetic studies using the cardiovascular gene-centric 50K SNP array. *Eur. J. Hum. Genet.*, 24, 106–112.

Shah, T. et al. (2013) Population genomics of cardiometabolic traits: design of the University College London-London School of Hygiene and Tropical Medicine-Edinburgh-Bristol (UCLEB) Consortium. *PLoS One*, 8, e71345.

Shi, J. et al. (2014) Characterizing the genetic basis of methylome diversity in histologically normal human lung tissue. *Nat. Commun.*, 5, 3365.

Shihab, H.A. et al. (2015) An integrative approach to predicting the functional effects of non-coding and coding sequence variation. *Bioinformatics*, 31, 1336–43.

Spain, S., and Barrett, J. (2015) Strategies for fine mapping complex traits. *Hum. Mol. Genet.*, 42, 1001–1006.

Teslovich, T.M. et al. (2010) Biological, clinical and population relevance of 95 loci for blood lipids. *Nature*, 466, 707–713.

UK10K consortium. (2015) The UK10K project identifies rare variants in health and disease. *Nature*, 526, 82–90.

Wood, A.R. et al. (2014) Defining the role of common variation in the genomic and biological architecture of adult human height. *Nat. Genet.*, 46, 1173–1186.

Yang, J. et al. (2012) Conditional and joint multiple-SNP analysis of GWAS summary statistics identifies additional variants influencing complex traits. *Nat. Genet.*, 44, 369–375.

Zheng, J. et al. (2013) Sequential sentinel SNP Regional Association Plots (SSSRAP): an approach for testing independence of SNP association signals using meta-analysis data. *Ann. Hum. Genet.*, 77, 67–79.