Research Article

Study of a New Chaotic Dynamical System and Its Usage in a Novel Pseudorandom Bit Generator

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A new chaotic discrete dynamical system, built on trigonometric functions, is proposed. With intent to use this system within cryptographic applications, we proved with the aid of specific tools from chaos theory (e.g., Lyapunov exponent, attractor’s fractal dimension, and Kolmogorov-Smirnov test) and statistics (e.g., NIST suite of tests) that the newly proposed dynamical system has a chaotic behavior, for a large parameter’s value space, and very good statistical properties, respectively. Further, the proposed chaotic dynamical system is used, in conjunction with a binary operation, in the designing of a new pseudorandom bit generator (PRBG) model. The PRBG is subjected, by turns, to an assessment of statistical properties. Theoretical and practical arguments, rounded by good statistical results, confirm viability of the proposed chaotic dynamical system and newly designed PRBG, recommending them for usage within cryptographic applications.

1. Introduction

Nowadays, more and more, it appears that skilful genesis of chaos turns out to be a key issue in many technological application fields such as engineering, medicine, communications, information storage, and, with particular importance, cryptography [1–6].

Designing of dynamical systems, intended to be used as base of cryptosystems, must be done so as to ensure the use of a set of associated control parameters’ values that leads to chaos [7–11]. Moreover, the ergodic [12–15] and randomness properties [16–19] must be confirmed, as a certainty of high security level of the chaotic dynamical system.

Since 1963, when Lorenz found the first chaotic attractor in a three-dimensional autonomous system while studying atmospheric convection [20], chaotification became a very attractive subject, leading to the development of new chaotic dynamic systems, for example, [21–25], and new chaos-based PRBGs, for example, [26–30], whose properties have been analyzed extensively and thoroughly in research or review articles and books, for example, [31–36].

Motivated by the extent of previous work, the present paper aims to present a new chaotic discrete dynamical system which, furthermore, may be included in the wide family of PRBGs through a simple, interesting, and yet complex new PRBG model, based on binary operation.

The rest of this paper is organized as follows. Section 2 presents the design of newly proposed dynamical system, including its chaotic behavior assessment (as a first step in system’s evaluation process to establish its suitability within any cryptographic application). Section 3 showcases the detailed and comprehensive randomness’ testing process of sequences generated by the new chaotic dynamical system (as a second step in system’s evaluation process to establish its suitability within any cryptographic application). Section 4 presents the designing of a new PRBG scheme (based on binary operation, which uses previously designed and tested chaotic dynamical system), including the results of analysis performed using NIST suite in order to test the randomness and uniformity of values generated by the new PRBG. Finally, Section 5 concludes the work carried out.
2. Proposed Discrete Dynamical System and Its Basic Properties

Newly dynamical system introduced in paper uses (1) [37, 38] as model for chaos generation. Here, whilst $f$ represents a periodic real map (selected so as to ensure a large phase space), $h$ represents a bounded real map (which, by an appropriate selection, restricts the phase space to a closed interval in which the dynamical system has good chaotic properties):

$$x_{n+1} = h(f(x_n)).$$  \hspace{1cm} (1)

Therefore, the newly proposed one-dimensional discreet dynamic system, which is defined with respect to form (1), is given by (2) or (and), in a more detailed design, by (3). Here, $r$ represents the control parameter of the resulted chaotic map, while $\arctg$ (i.e., arctangent function), and $\ctg$ (i.e., cotangent function), respectively, were chosen with respect to the above affirmations (i.e., the first one restricts the phase space to a close interval in which the dynamical system has good chaotic properties, while the second one ensures a large phase space):

$$x_{n+1} = f_p(x_n),$$  \hspace{1cm} (2)

$$f_p : [-1, 1] \rightarrow [-1, 1], \quad f_p(x) = \frac{2}{\pi} \arctg(\ctg(rx)).$$  \hspace{1cm} (3)

In the following, dynamical behavior of newly proposed chaotic system is investigated, by both theoretical analysis and numerical simulation (e.g., by means of Lyapunov exponent, attractor’s geometric shape and fractal structure, and system’s ergodicity, i.e., Kolmogorov-Smirnov tests, etc.).

2.1. Sensitivity Level to Initial Conditions. The behavior of the proposed discrete dynamic system (3), in terms of its evolution over time domain, depends on both the control parameter $r$ and the initial condition $x_0$. First of all, we propose stability analysis of fixed points in order to assess system’s sensitivity level to initial conditions.

$f_p$ map’s fixed points are given by (4), and, according to the theorem of fixed points [39], $x_k$ points are attractors if condition (5) is fulfilled, that is, if the control parameter $r$ meets condition (6). Thus, taking into account the fact that $f_p$ map is defined on the interval $[-1, 1]$, there exists only one fixed point, that is, (7), which is also an attractor:

$$t_k = \frac{\pi (2k + 1)}{2r + \pi}, \quad k \in \mathbb{Z},$$  \hspace{1cm} (4)

$$|f_p'(t_k)| < 1, \quad k \in \mathbb{Z},$$  \hspace{1cm} (5)

$$r < \frac{\pi}{2},$$  \hspace{1cm} (6)

$$t_0 = \frac{\pi}{2r + \pi}.$$  \hspace{1cm} (7)

From the above equations, it can be noticed that for any $r \in (0, \pi/2)$, all trajectories that start at initial point $x_0$ converge, in time, to the attractor point $t_0$; for any $r > \pi/2$, fixed point $t_0$ loses its stability, and other instable fixed points appear. The aforementioned statements are also substantiated by the bifurcation diagram, namely, the one showcased in Figure 1(a), which emphasizes the stability of $f_p$’s fixed points (i.e., for any values of parameter $r$, close to $\pi/2$, $f_p$ has an instable behavior, and for any values of the parameter $r$, higher than $\pi/2$, the map enters into a complete chaotic regime). The road to chaos of the $f_p$ map, with $r > \pi/2$, is not achieved through doubling process of the period, specific to some chaotic maps, but is induced by existence of a dense set of periodic orbits, whose periods are in the $[-1, 1]$ interval.

Another tool used to assess $f_p$’s sensitivity to the initial conditions is the Lyapunov exponent (8) and, taking into consideration the fact that orbit $\{x_1, x_2, \ldots, x_n\}$ is chaotic if this exponent is positive, (9) is derived (which is equivalent to $r > \pi/2$, meaning that for any $r \in (\pi/2, 10)f_p$’s orbits are chaotic). Lyapunov exponent, numerically computed [40] with respect to the parameter $r$ within $(0, 10)$ interval, is shown in Figure 1(b). Consider

$$\lambda = \lim_{n \to \infty} \frac{1}{n} \sum_{i=1}^{n} \ln |f_p'(x_i)|$$

$$= \lim_{n \to \infty} \frac{1}{n} \sum_{i=1}^{n} \ln \left| \frac{-2r}{\pi \sin^2 rx (1 + \ctg^2 rx)} \right|$$
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2.2. Attractor’s Geometric Shape and Fractal Structure. Analysis of dynamical system’s attractor’s shape can provide meaningful information about system behavior in time, for certain values of its parameters. The attractor of a dynamical system with a periodic behavior has a regular shape, while the one corresponding to a chaotic dynamical system has a complex structure, of fractal type, called strange attractor [39].

Figure 2 showcases $f_p$’s attractor, for $r = 7$, where it can be observed that its shape is irregular, of fractal type, and complex shaped, in comparison with ones of, for example, tent or logistic maps (i.e., whose attractors exhibit a regular shape, i.e., triangle, resp., hyperbola shaped).

Fractal structure of an attractor is indicated by a fractional value of its fractal dimension, which is a ratio that provides a statistical index of complexity comparing how in detail a pattern changes with the scale at which it is measured or, alternatively, by a measure of the space-filling capacity of a pattern, telling how a fractal scale is different than the space in which it is embedded. There are several types of fractal dimensions, which can be theoretically and empirically estimated, such as Hausdorff dimension, Minkowski-Bouligand dimension, box-counting dimension, information dimension, and correlation dimension [41–44]. Using plots from Figure 3, we established that the attractor of the $f_p$ map has a box-counting dimension $D_b = 0.97863$ and a correlation dimension $D_c = 0.97064$. Fractional values, of fractal dimensions previously estimated, allow us to conclude that the proposed map has a strange attractor which, by turns, indicates a chaotic behavior.

2.3. System’s Ergodicity. In this subsection, using Birkhoff’s theorem [5, 45] in conjunction with Kolmogorov-Smirnov test [46, 47], we intend to prove that the proposed dynamical system is ergodic for $r > \pi/2$ (i.e., long-term behavior of $f_p$’s orbits is independent from the initial condition, and thus it may be subjected to a battery of statistical tests).

The Kolmogorov-Smirnov test is applied on two amounts of independent data $(x_1, x_2, \ldots, x_n)$ and $(y_1, y_2, \ldots, y_n)$, corresponding to the measurements of two random variables $X$ and $Y$. Random variable $X$ is obtained through an $n$-times iteration of $f_p$ map, for a fixed parameter $r$ and for a fixed initial condition $x_0$ (namely, $r > \pi/2$ and $x_0 \in [-1, 1]$). Second random variable, that is, $Y$, is obtained by selecting, at time $k$, the values generated by $n$ orbits of the map, arising from $n$ initial seeding points (belonging to $[-1, 1]$ interval) and same $r$ parameter (previously fixed). Moment $k = 100$ is chosen from $f_p$’s stationary zone, previously established using Kolmogorov-Smirnov test, as described in [46, 47].

Due to the fact that the random values $X$ and $Y$ correspond to time average of $f_p$ and space average, respectively, the purpose of the test is to establish if the two experimental data sets derive from populations with the same distribution or not, with respect to Birkhoff’s theorem. The analysis is based on distribution functions $F_X$ and $F_Y$, associated with the experimental independent data sets $X$ and $Y$.

Kolmogorov-Smirnov test is applied as follows:

1. the maximum absolute difference between the two distribution functions $\delta$ is computed:
   \[
   \delta = \max_u \left| F_X (u) - F_Y (u) \right| ;
   \]  

2. for a significance level $\alpha$, $\Delta_\alpha$ is computed ($\alpha$ representing probability law’s quantile, for the random value $\Delta$); that is, $P(\Delta > \Delta_\alpha) = \alpha$,
   \[
   \Delta_\alpha \equiv \sqrt{\frac{nm}{nm}} \left[ \frac{1}{2} \ln \frac{2}{\alpha} \right] ;
   \]

3. in case of $\delta \leq \Delta_\alpha$, the hypothesis $H_0$ is accepted (i.e., the two random variables $X$ and $Y$ have the same probability law; in other words, if the absolute maximum distance between the two distribution functions $F_X$ and $F_Y$ is lower than a certain accepted value $\Delta_\alpha$, then it will be decided that the two random variables $X$ and $Y$ have the same probability law); otherwise the test rejects $H_0$ hypothesis (i.e., for the chosen level, the two sets of experimental data come from random values with different probability laws).

Kolmogorov-Smirnov test was performed over a sequence of $n = m = 100,000$ samples, with a significance level set to $\alpha = 0.05$; the decision regarding system’s ergodicity is based on Monte Carlo analysis (i.e., evaluating Kolmogorov-Smirnov test’s ability to accept bad data as good data). The above experiment was repeated 500 times; at the end of each round, $H_0$ hypothesis’s acceptance proportion (which belongs to $[0.93, 0.97]$ interval) was recorded.

Test’s overall results are summarized in Table 1. One can observe that in case of all values selected for $r$ parameter, with $r \in (\pi/2, 10)$, the acceptance proportion of $H_0$ hypothesis lies within the confidence interval. Thus, ergodicity of the proposed dynamical system is confirmed (over the entire interval of interest, for parameter $r$).
Based on numerical results previously obtained, using instruments from the chaos theory, we can conclude that $f_p$ map has a chaotic behavior, without intermittent scenarios, for any combination between the parameter $r \in (\pi/2, 10)$ and the initial seeding point $x_0 \in [-1, 1]$.

The chaotic behavior is a necessary but not sufficient condition to allow usage of the proposed dynamic system within cryptographic applications. System’s security level, against some statistical cryptanalytic attacks, is assessed after a statistical analysis of the randomness of values generated. There are several options available for analyzing randomness of a newly developed pseudorandom bit generator (PRBG), as it will be revealed in the following section.

### 3. Randomness Analysis of the Proposed Chaotic Discrete Dynamical System

In order to assess PRBG’s statistical properties (i.e., its true randomness and implicit suitability within cryptographic applications; see e.g., [48–52], etc.), different testing tools such as CrypTool and VRA (for basic statistical measures’ quantification), respectively NIST [53] and DIEHARD [54] standard tests batteries (for high end quantitative and qualitative assessment) were used. Operating methodology, for each of the above tests, and obtained results are presented and discussed in the following subsections.

Chaotic cryptography deals with real numbers, so, in order to proceed and apply the battery of the statistical tests aforementioned, we have to apply a computational method to transform a chaotic sequence of real numbers into a bitstream. The discretization method that we used consisted in the extraction of the fractional parts of the generated subunitary real numbers.

#### 3.1. CrypTool Analysis

CrypTool was used to compute the occurrence frequencies of any binary substring, composed of $n$ symbols (i.e., the $n$-grams), over the flow of bits generated with the proposed PRBG. For true random bitstreams, it is expected that each entry within the $n$-gram has the same probability of occurrence. The $n$-gram statistics were performed over 1,000 randomly chosen binary sequences, each sequence being generated using different initial seeding
points and of \( i = 1,000,000 \) bits in length; overall results are presented in Table 2.

It can be observed that the deviation from ideal value, of each \( n \)-gram's entry, is under 0.1%. Thus, PRBG's \( n \)-gram reports not only do not emphasize the dominant presence of any binary substring (i.e., in terms of frequency of use) [55, 56], but they also highlight a uniform system dynamics (i.e., in terms of the time evolution of \( f_p \)'s trajectories). Positive results obtained at this point guide us to perform the next statistical analysis.

### 3.2 Visual Recurrence Analysis

A RP (i.e., Recurrence Plot) holds important insights into the time evolution of \( f_p \)'s trajectories because typical patterns in RPs are linked to specific system behavior [57]. Yet, without proper settings of analysis' parameters, any RP is just a simple image, completely devoid of information. Thus, to obtain as much information, suitable embedding dimension and adequate time delay must be chosen. With the aid of AMI (i.e., Average Mutual Information) and FNN (i.e., False Nearest Neighbors) toolboxes (included in VRA's statistical test suite), these parameters can be correctly set to the optimal value [58]. AMI and FNN were performed on binary sequences, each of \( i = 1,000,000 \) bits in length (generated with 100 randomly chosen seeds); corresponding graphs are presented in Figures 4(a) and 4(b), respectively. Analyzing the two graphs, first AMI minimum and Optimal Global Embedding Dimension (OGED) are found: time delay \( d = 3 \) and embedding dimension \( m = 3 \). With these two parameters bitstream's RPI (i.e., Recurrence Plot Image) is computed, as shown in Figure 4(c).

Lack in clear patterns, within the RPI, indicates that consecutive samples in bitstream's structure are much far apart and uncorrelated. More than that, RPI's homogeneity along the major diagonal and its irregular distribution emphasizes a stationary, mostly stochastic behavior (i.e., intrinsically nondeterministic, nonintermittent, and sporadic), of the system that has generated the bitstream and, namely, a true random process (i.e., random binary strings).

VRA, through its embedded RQA (i.e., Recurrence Quantification Analysis) tool, also provides other additional measures (e.g., entropy, mean, percentage of recurrence and of determinism, etc.); some of them, the most important ones, are quantified in Table 3. One can notice that all the measures have values close to ideal [59].

Despite the fact that skewness has a negative value (i.e., indicating that the tail on the left side of the probability density function is longer than the right side and the bulk of the values lie to the right of the mean), being close to zero indicates that the values are relatively evenly distributed on both sides of the mean, typically (but not necessarily) implying a symmetric distribution [60]. At the same time, Kurtosis's high-level and negative value denotes a platykurtic distribution (i.e., data set with flatter peak around its mean, which causes thin tails within the distribution and low level of data fluctuation) [61].

Good general statistical properties revealed with the aid of VRA (either visually—evaluation of RPI's structural properties or through different specific measures evaluation—RQA), highlights randomness of bitstreams generated using the proposed PRBG function, thus allowing advancement to other statistical test suites.

### 3.3 NIST Statistical Testing

For the numerical experimentations of the proposed pseudorandom bit generator, we have generated \( m = 2,000 \) different binary sequences from 500 randomly chosen seeds, each sequence having a length of \( n = 1,000,000 \) bits, and we have computed the \( p \)-value corresponding to each sequence for all the 15 tests of the NIST suite. The significance level of each test in NIST is set to 1%, which means that 99% of test samples pass the tests if the random numbers are truly random. The acceptance region

| \( n \)-gram's order | Substring | Frequency (%) | \(|\Delta|\) |
|---------------------|-----------|---------------|----------|
| Histogram (\( n = 1 \)) | 0         | 50.0734       | 0.0734   |
|                     | 1         | 49.9266       | 0.0734   |
| Digram (\( n = 2 \))  | 00        | 24.9605       | 0.0395   |
|                     | 01        | 24.9660       | 0.0340   |
|                     | 10        | 24.9660       | 0.0340   |
|                     | 11        | 25.1074       | 0.1074   |
| Trigram (\( n = 3 \)) | 000       | 12.4899       | 0.0101   |
|                     | 001       | 12.4706       | 0.0294   |
|                     | 010       | 12.4293       | 0.0707   |
|                     | 011       | 12.5367       | 0.0367   |
|                     | 100       | 12.4705       | 0.0295   |
|                     | 101       | 12.4954       | 0.0046   |
|                     | 110       | 12.5367       | 0.0367   |
|                     | 111       | 12.5707       | 0.0707   |
| 4-gram (\( n = 4 \))  | 0000      | 6.2240        | 0.0260   |
|                     | 0001      | 6.2659        | 0.0159   |
|                     | 0010      | 6.2015        | 0.0485   |
|                     | 0011      | 6.2691        | 0.0191   |
|                     | 0100      | 6.2085        | 0.0415   |
|                     | 0101      | 6.2208        | 0.0292   |
|                     | 0110      | 6.2568        | 0.0068   |
|                     | 0111      | 6.2799        | 0.0299   |
|                     | 1000      | 6.2658        | 0.0158   |
|                     | 1001      | 6.2047        | 0.0453   |
|                     | 1010      | 6.2278        | 0.0222   |
|                     | 1011      | 6.2676        | 0.0176   |
|                     | 1100      | 6.2620        | 0.0120   |
|                     | 1101      | 6.2746        | 0.0246   |
|                     | 1110      | 6.2799        | 0.0299   |
|                     | 1111      | 6.2908        | 0.0408   |
| 12-gram (\( n = 12 \)) | 000000000000 | 0.0265 | 0.0020 |
|                     | —         | ≈0.0244       | ≈0   |
|                     | 111111111111 | 0.0284 | 0.0039 |
of the passing ratio is given by (12), where $m$ represents the number of samples tested and $p = 1 - \alpha$ is the probability of passing each test:

$$\left[ p - 3 \sqrt{\frac{p(1-p)}{m}}, p + 3 \sqrt{\frac{p(1-p)}{m}} \right].$$ (12)

For $m = 2000$ and the probability $p = 0.99$ (corresponding to the significance level $\alpha = 0.01$), we obtained the confidence interval $[0.983, 0.996]$. In the second column of Table 4, we have summarized the results obtained after applying nonparameterized and parameterized tests of the NIST suite on the binary sequences produced by the proposed pseudorandom bit generator. The computed proportion for each test lies inside the confidence interval. Hence, the tested binary sequences generated by the proposed PRBG are random with respect to all tests of NIST suite. If tested sequences are truly random, then $p$-values are expected to appear uniform in the interval $[0,1)$. NIST recommends to apply the $\chi^2$-test in which the interval $[0,1)$ is divided into 10 subintervals. Defining $F_i$ as number of occurrences of the $p$-value in $i$th interval, then the $\chi^2$ statistic is (13). NIST recommends setting its significance level to 0.01%, so the acceptance region of statistics has the value $\chi^2 \leq 33.72$. A $p$-value that corresponds to uniformity of $p$-values is calculated as in (14), so it must be greater than 0.0001.
Table 4: NIST tests’ results.

| Test name                | Passing ratio of the test | Uniformity p-value | Test result |
|--------------------------|---------------------------|--------------------|-------------|
| Frequency                | 0.992                     | 0.602803           | PASSED      |
| Block frequency          | 0.990                     | 0.748891           | PASSED      |
| Cumulative sums          | 0.991                     | 0.090388           | PASSED      |
| Runs                     | 0.990                     | 0.939005           | PASSED      |
| Longest run              | 0.989                     | 0.592443           | PASSED      |
| Rank                     | 0.991                     | 0.840367           | PASSED      |
| FFT                      | 0.989                     | 0.242363           | PASSED      |
| Nonoverlapping template  | 0.983                     | 0.761719           | PASSED      |
| Overlapping template     | 0.983                     | 0.230755           | PASSED      |
| Universal                | 0.987                     | 0.050629           | PASSED      |
| Approximate entropy      | 0.988                     | 0.959347           | PASSED      |
| Random excursions        | 0.987                     | 0.614382           | PASSED      |
| Random excursions variant| 0.984                     | 0.830939           | PASSED      |
| Serial                   | 0.986                     | 0.209392           | PASSED      |
| Linear complexity        | 0.989                     | 0.764655           | PASSED      |

to ensure that the p-values could be considered uniformly distributed. The results from the third column of Table 4 lead us to the conclusion that p-values, for each statistical test, are uniformly distributed:

\[
\chi^2 = \sum_{i=1}^{10} \frac{(F_i - m/10)^2}{(m/10)},
\]

\[
\text{igamc}(9/2, \chi^2/2).
\]

The method to calculate the passing ratio of total test and the uniformity p-value of total test samples follows the same methodology described above. In this case, we considered the number of samples \( m = 30,000 \), so the acceptance region is \([0.988, 0.992]\). For the passing ratio of the total test we obtained the value 0.988, and the p-value corresponding to the uniformity of p-values from the total test was 0.294808, so the proposed map has perfect cryptographic properties.

4. Proposed PRBG and Its Statistical Testing

Most chaos-based PRNGs (and, implicitly, their subsequent PRBGs) are based on a single chaotic system (e.g., [25–29, 52]) and generate the random numbers (resp., the bitstreams) directly from its orbit. These types of PRNGs/PRBGs are potentially insecure, since the output values may expose some information about their underlying chaotic system [62]. To overcome this difficulty, a series of pseudorandom number (bit) generators based on a couple of chaotic system have been proposed (e.g., [63–67]). In the following, we present a novel PRNG/PRBG model, based on two chaotic maps coupled using a binary operation.

We consider two one-dimensional chaotic maps (e.g., as the previously designed model, i.e., (3)) defined as follows:

\[
f_1 : [-1, 1] \rightarrow [-1, 1], \quad f_1(x, r_1) = \frac{2}{\pi} \arctg(\tan(r_1 x)),
\]

\[
f_2 : [-1, 1] \rightarrow [-1, 1], \quad f_2(x, r_2) = \frac{2}{\pi} \arctg(\tan(r_2 x)),
\]

(15)

where \( x_1^0, x_2^0 \) are the initial conditions, \( r_1, r_2 \) are the control parameters, and \( x_1^i, x_2^i \) are the two orbits obtained by recurrences \( x_{i+1}^1 = f_1(x_i^1, r_1), x_{i+1}^2 = f_2(x_i^2, r_2), \) for any \( i \in \{0, 1, 2, \ldots\} \).

Also, we consider the binary operation given by the formula

\[
a \ast b = \frac{a + b}{1 - a \cdot b},
\]

(16)

where \( a, b \in [-1, 1] \).

The output \( \{y_0, y_1, y_2, \ldots\} \) of the proposed PRNG is obtained applying the binary operation (16) to the chaotic maps \( f_1 \) and \( f_2 \) defined by (15); thus,

\[
y_i = f_1(x_i^1, r_1) \ast f_2(x_i^2, r_2) = \frac{f_1(x_i^1, r_1) + f_2(x_i^2, r_2)}{1 - f_1(x_i^1, r_1) \cdot f_2(x_i^2, r_2)}
\]

(17)

for any \( i \in \{0, 1, 2, \ldots\} \) and \((x_0^0, r_1, x_0^2, r_2) \in \{-1, 1\} \times [1, 10] \times \{-1, 1\} \times [1, 10] \) is the seed of the proposed PRNG.

The real numbers obtained using the proposed PRNG were discretized extracting their fractional parts in order to apply the NIST statistical tests. For the numerical experiments on the proposed pseudorandom numbers generator, we have generated 2,000 different binary sequences (sample size \( m = 2,000 \)). Each sequence, with one million bits in length, has been generated from a randomly chosen seed.
The computed proportion for each test lies inside the confidence interval; hence, the tested binary sequences generated by the proposed PRBG are random with respect to all the 16 tests of NIST suite.

5. Conclusions

Development of new chaotic dynamic systems, which meet the current demands of security, is a present research direction in the field of cryptography. The main objective is to obtain a large key space, induced by the control parameter and (or) initial conditions, for which the dynamic system is in chaotic regime, is ergodic, and has a uniform distribution of the values generated.

With respect to the aforementioned ideas, in this paper we have designed a new one-dimensional chaotic dynamic system that meets these requirements.

Moreover, a larger key space than the one of the known chaotic maps (e.g., logistic, tent, Hénon, etc.) was achieved, and, despite the fact that the implementation of trigonometric maps is little slower than the ones of other kinds of maps (e.g., polynomial, exponential, etc.), we consider that the advantage of a larger key space induced by their usage is a good compromise (i.e., a win-win situation).

Using specific mathematical and numerical tools from chaos theory and statistics, we proved that the proposed chaotic dynamic system has very good cryptographic properties. The proposed map was used in a new innovative way to design a new PRNG/PRBG model, based on a well-known binary operation.

We have performed an exhaustive testing process of the randomness of the generated binary sequences using the NIST suite to prove the viability of the proposed PRNG/PRBG.

Conflict of Interests

The authors declare that there is no conflict of interests regarding the publication of this paper.

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