Abstract

Stochastic simulators are an indispensable tool in many branches of science. Often based on first principles, they deliver a series of samples whose distribution implicitly defines a probability measure to describe the phenomena of interest. However, the fidelity of these simulators is not always sufficient for all scientific purposes, necessitating the construction of ad-hoc corrections to “calibrate” the simulation and ensure that its output is a faithful representation of reality. In this paper, we leverage methods from transportation theory to construct such corrections in a systematic way. We use a neural network to compute minimal modifications to the individual samples produced by the simulator such that the resulting distribution becomes properly calibrated. We illustrate the method and its benefits in the context of experimental particle physics, where the need for calibrated stochastic simulators is particularly pronounced.

1. Introduction

Many of the physical sciences heavily rely on complex simulations to describe processes and phenomena of interest. One important category of simulation programs, relevant to this paper, is formed by stochastic simulators [1]. They model systems or processes that involve elements of randomness, i.e. whose behaviour is not uniquely determined by the set of initial conditions.

The challenges associated with the construction of such simulators are nicely illustrated by an example from experimental particle physics. In scattering experiments, collisions between energetic particles can access information about the fundamental physical dynamics at very short length scales. As probes of quantum mechanical—and hence probabilistic—processes, only the distribution of a large number of these scattering “events” carries meaningful information.

To exploit the experimentally collected data to the fullest extent, it is important to have access to accurate predictions of the expected event distribution. Such “event generators”
must incorporate a large number of physical effects occurring at very different energy scales, requiring methods ranging from perturbative calculations to parametrised, phenomenological models. The resulting simulation models generally have a very large number of free parameters. Many of these parameters describe well-understood physical processes (typically at low energy scales), which are not of interest per se, but must be included in order to faithfully represent physical reality. Only a small subset directly concerns the fundamental physics of high-energy scattering reactions, and are thus of most interest to the experimenter.

Before being applied to the study of novel physical phenomena, it is vital to ensure that the simulation can accurately model all known contributing effects. For the purposes of this “calibration”, the output of the simulator is compared to a suitable calibration dataset, usually taken from experiment. The simulation model is then adjusted such that discrepancies are largely removed (While the word “calibration” may carry many meanings, depending on context, in this paper we take it to mean the correction of a simulation to reproduce some aspect of observed data). Often, the causes underlying these differences cannot be interpreted in terms of the original simulation parameters, or it is simply not feasible to do so. Therefore, the applied calibration does not typically take the form of a well-motivated change of the simulation model or its parameter values, but is rather an ad-hoc modification of its output.

In this paper, we present a very general and flexible way for the calibration of such a stochastic simulator, and apply this technique to situations that are typical for experimental particle physics. In Section 2, we give a brief overview of the simulation models and calibration techniques that are used in this field. Section 3 then formalises the task of calibrating a given simulator. Section 4 introduces our proposed method by rephrasing this calibration problem in the language of transportation theory, and Section 5 summarises aspects of its practical implementation. Finally, Section 6 documents several typical example applications.

Reading guide. Section 2, on applications and calibration techniques in particle physics, has been written specifically for practitioners of the subject, but can be read independently of the remainder of the paper. The main body of the manuscript, Sections 3 to 5, presents our proposed calibration method in a very general way, and should be accessible as well to readers who are not familiar with particle physics. The description of the example applications shown in Section 6 while again inspired by situations from particle physics, should likewise be accessible.

2. Calibrated simulators in experimental particle physics

Particle collider experiments make use of sophisticated collision and detector simulations that have been refined over many decades [2, 3, 4, 5]. While these simulations have been carefully tuned and match observed data remarkably well, the large datasets coming from machines such as the LHC, for instance, lead to extremely demanding environments. It is common, therefore, to calibrate various aspects of these simulation models as part of the data analysis to remove residual simulation deficiencies. Quantities in need of calibration often feed rather directly into the final analysis results, underlining the importance of well-behaved calibration techniques.
For example, corrections to the average muon and electron energy responses in relevant detectors are derived using decays of $Z$ bosons to charged leptons \[6, 7, 8, 9\], taking advantage of precision measurements of the $Z$ boson pole mass from previous experiments \[10\]. Similar modifications are derived for the first two moments of the jet energy response using momentum-balancing techniques at hadron colliders \[11, 12\]. The calibration is applied by modifying the respective quantity or object in-situ, as shown in Figure \[1(a)\].

![Figure 1: In (a) the calibration replaces a jet with measured energy $E$ with a jet with a different energy $E'$ in order to correct the jet energy scale and resolution. In (b) scale-factors are used to ensure proper modelling of the distribution of a discriminant $D$.](image)

The LHC experiments also make heavy use of high-level quantities with less-obvious physical meanings than “energy” and “mass”. For example, high-performance discriminators, which are built from artificial neural networks or boosted decision trees and are trained to identify “interesting” objects in collision events based on many input signals from the detector, e.g. for $b$-tagging \[13, 14\], are ubiquitous in modern collider data analyses. These discriminators involve complex calculations with dozens or hundreds of input features, and may in turn appear rather opaque. Correcting the collision simulation for these intricate observables can be a challenge because the detector response is not very strongly constrained by the physicist’s prior knowledge or intuition—simply correcting the mean or width of the output distribution is very likely not sufficient, as it is for a distribution that is approximately gaussian like the lepton or jet energy response.

One common technique for the calibration of such general distributions involves calculating binned density ratios between collision data and predictions from the simulation for the distribution of interest, as illustrated in Figure \[1(b)\]. This has proven very useful e.g. in the case of $b$-jet tagging \[13, 14\]. These density ratios are commonly known as “scale-factors”; in the calibration procedure, they act as additional weights associated to the calibrated object. Scale-factors are derived in well-understood control samples and may be parametrised by one or more other observables; the transverse momentum $p_T$ or the pseudorapidity $\eta$ are common scale-factor parameters at a hadron collider.
Density ratios built from binned distributions do come with some drawbacks, however. Bins with no predicted events cannot be corrected, and the scale-factor weights in general change the total number of predicted events when applied in regions that do not have the same pre-calibration distribution in simulation. In addition, binned density estimates generalise very poorly to multidimensional settings, which makes this technique unsuitable for calibrations where dependencies on many variables need to be considered. Recent developments derive high-dimensional, un-binned scale factors using boosted decision trees [15], artificial neural networks [16], and adversarial neural networks [17], although the possibility of changing the total number of predicted events by applying these weights remains.

In the following, we lay out an approach based on transportation theory that does not suffer from these shortcomings.

3. Formalisation of the problem

We now make the notion of “calibrating” a stochastic simulator more precise, and set up the notation that will be used in the following. For the purposes of this paper, a stochastic simulator is an arbitrary computational program that implicitly defines a probability distribution \( p(x|\theta) \) by generating samples (or “events”) \( x \in X \) that are distributed according to this distribution, i.e. \( x \sim p(x|\theta) \). The vector \( \theta \) labels the set of free parameters of the simulation, e.g. physical constants. (The simulation can also have discrete parameters, which e.g. select a specific model from a set of alternatives.) We shall not impose any restrictions on the form and content of the samples \( x \). In general, \( x \) is a high-dimensional vector whose components fully describe the final state of the simulation. In certain cases, it can be convenient to allow the simulator to deliver weighted samples, i.e. to produce pairs \((x, w)\), where \( w \) takes the role of an event weight.

A set of generated samples, \( X = \{(x_i, w_i)\}_{i=1}^N \) allows expectation values of arbitrary observables to be estimated, \( \langle O \rangle \approx \sum_{i=1}^N w_i O(x_i) / \sum_{i=1}^N w_i \). If required, density estimation techniques (such as histograms) give access to an explicit estimate of the underlying probability distribution \( p(x|\theta) \).

To calibrate the simulator, the simulated probability distribution \( p(x|\theta) \) is compared to the distribution \( q(y) \), \( y \in Y \) the simulator is intended to sample from. That is, \( q(y) \) describes the actual behaviour of a physical system, and not merely its model. In practice, \( p(x|\theta) \) is usually very close (but not identical) to \( q(y) \) for suitable simulation parameters \( \theta \). As these densities are intractable and not directly accessible, the calibration is performed by comparing the simulated dataset \( \mathcal{X} \) to the calibration dataset \( \mathcal{Y} = \{(y_j, w_j)\}_{j=1}^M \) drawn from \( q \). In many cases \( \mathcal{Y} \) is obtained experimentally, which usually results in events \( y_j \) that carry unit weight.

To eliminate any residual differences between \( p(x|\theta) \) and \( q(y) \) (and thus also between \( \mathcal{X} \) and \( \mathcal{Y} \)), a calibration procedure is applied to the output of the uncalibrated simulator, i.e. the dataset \( \mathcal{X} \). Given an event \((x_i, w_i) \in \mathcal{X} \), any admissible calibration procedure thus generates a modified event \((x'_i, w'_i) \). The collection of these events, \( \mathcal{X}' = \{(x'_i, w'_i)\}_{i=1}^N \), is then said to come from the calibrated simulator, which samples from the calibrated distribution \( p'(x'|\theta) \). For this procedure to be successful, \( p'(x'|\theta) \) should be indistinguishable from \( q(y) \) for the
purposes of computing arbitrary expectation values and estimating densities. We denote this equivalence as \( p'(x'|\theta) \equiv q(y) \). (In practice, of course, the quality of the calibration may only be evaluated based on the equivalence of the corresponding collections of samples \( \mathcal{X}' \) and \( \mathcal{Y} \).) A precise formulation of this requirement in the context of our proposed calibration method follows below in Section 4.

4. Transportation theory for calibrations

The calibration strategies mentioned in Section 2 belong to two distinct categories. Scale factors solely affect modifications to the event weight, i.e. implement calibration procedures of the type \( (x_i, w_i) \rightarrow (x_i, w'_i) \). Modifications of the to-be-calibrated quantities themselves (such as the jet energy) manifestly do not change the weight, and thus correspond to a calibration step of the form \( (x_i, w_i) \rightarrow (x'_i, w_i) \). In a typical application in particle physics, techniques of both kinds can be present simultaneously to correct various different aspects of the simulation.

In this section we focus on the latter case, and outline a systematic method to construct such a calibration for very general situations. As we shall see, in various special cases, this technique reduces to prescriptions that are already well-known to practitioners, and thus motivates and formally justifies them.

A general, weight-agnostic calibration method brings with it a number of immediate conceptual benefits. First, modifying the event directly leads to greater flexibility in the application of the calibration: the calibration can easily be restricted to only affect certain components of \( x \) (e.g. those that are known to be poorly modelled) while ensuring that components for which the simulator is thought to be accurate are not modified at all. Second, the normalisation of the dataset is manifestly preserved by the calibration, while this need not always be the case for weight-based techniques. Finally, no additional weight variance is induced by the calibration and the statistical power of the available samples is not diluted.

Here and in the following, we consider only the case where all components of the event vector \( x \) are continuous, and the corresponding probability densities are sufficiently smooth. This restriction is not strictly needed, and our method can be extended to also apply to discrete probability distributions, albeit at the expense of added technical complexity.

4.1. Description of the method

The goal thus becomes to compute a deterministic correction \( x' = T(x; \theta) \) for each event such that the simulator becomes properly calibrated, i.e. \( p'(x'|\theta) \equiv q(y) \). Note that, since the uncalibrated distribution \( p(x|\theta) \) depends on the parameters \( \theta \), but the calibration target \( q(y) \) does not, also the function \( T \) implementing the calibration becomes dependent on \( \theta \). The function \( T \) need not be unique, i.e. there can exist nonequivalent solutions which lead to identical calibrated distributions \( p'(x'|\theta) \). A simple example is shown in Figure 2.

However, the uncalibrated simulator already contains a significant amount of domain knowledge. Naturally, the calibration should not negate this structure, but merely fine-tune the simulator output. Therefore, one should give preference to the function \( T(x; \theta) \)
Figure 2: Events $x = (x^1, x^2)$ produced by the uncalibrated simulator with simulation parameters $\theta_0$ are uniformly distributed in the inside of an axis-aligned square (blue), while a properly calibrated simulator uniformly populates with events $x' = (x'^1, x'^2)$ a square which is rotated by $\pi/2$ (green). The function $T^A(x; \theta_0)$ in (a) achieves a proper calibration by actually rotating the distribution, while $T^B(x; \theta_0)$ in (b) performs a more complicated transformation that leads to the same result.

that modifies the simulator to the minimal degree possible, while still achieving a proper calibration.

That is, the preferred solution to the calibration problem is the one with minimal expected modifications of the simulator,

$$\hat{T}(x; \theta) = \arg \inf_{T(x; \theta); p'(x'|\theta) \equiv q(y)} \int dx p(x|\theta) c[x, T(x; \theta)],$$

(1)

where the minimum is taken over all functions $T$ which successfully calibrate the simulator. Written explicitly, the requirement $p'(x'|\theta) \equiv q(y)$ asserts that

$$\int_{T^{-1}(B; \theta)} dx p(x|\theta) = \int_B dy q(y), \quad \forall B \in Y,$$

i.e. that the calibrated distribution $p'(x'|\theta)$ and $q(y)$ allocate equal probability mass to arbitrary regions $B \in Y$. The “cost function” $c[x, T(x; \theta)]$ provides a measure for the degree to which an event is modified by the calibration procedure. This distance measure should be chosen in accordance with the physics content of the observables that are being calibrated. It forms an important input to the method that allows the user to encode prior knowledge and steer the properties of the solutions of Equation 1. For quantities that are inherently unbounded (e.g. energy differences), the Euclidean distance might be a reasonable choice. In general, more complicated distance measures might be necessary to fully capture the content of the participating objects (angular variables are a common example). This aspect is illustrated below in Section 4.2. It’s interesting to note that similar ideas have been studied recently in high energy physics to define metrics between collider events [18, 19].

The optimisation problem in Equation 1 has been intensely studied for more than a century in the context of transportation theory, where its solution is interpreted as the
optimal way to transport a commodity between locations of production and locations of
demand. In this context, it is also known as the “Monge optimal transport problem” \[20\].

The proper analysis of the problem in Equation 1 requires levels of mathematical rigour
that might obstruct its implementation from a practitioner’s perspective, which we prefer to
take here. We thus proceed by briefly sketching some relevant aspects on an intuitive level
and refer to the mathematical literature (see references below) for all details.

4.2. Towards a solution for the quadratic cost

The characteristics of the solution of Equation 1 turn out to strongly depend on the form
of the cost function \( c[x,x'] \). It is easy to see that for the Euclidean cost \( c[x,x'] = \|x - x'\| \),
many equally “optimal” solutions are possible (imagine shifting a row of books to the right
by one: one can either move the entire collection of books by one place to the right, or pick
the leftmost book and place it at the far right \[21\]).

A special role is played by the quadratic cost \( c[x,x'] = \frac{1}{2} \|x - x'\|^2 \). For this case, the
optimal solution to Equation 1 is unique, and can be constructed as the gradient of a
scalar field, which plays the role of a “potential”. Note that this property already excludes
the function \( T^A(x;\theta_0) \) in Figure 2(a) as the optimal solution of the situation shown there:
rotations produce nonvanishing curl, while gradient fields are always curl-free. It turns out
that the function \( T^B(x;\theta_0) \) in Figure 2(b) does stem from a potential, and that, moreover,
it is indeed the solution that minimises the expectation of the quadratic cost. The event
modifications resulting from \( T^B(x;\theta_0) \) contain a significant radial component. However,
if the radial distribution of events is known to be well-modelled by the simulation, any
such radial transport should be avoided. In this example, one can use a distance measure
that—unlike the quadratic cost—differentiates between modifications along the radial and
azimuthal directions and that can thus suppress any radial transport\[4\]. This provides a way
to effectively channel corrections to the simulation so that they are compatible with prior
expectations.

To construct the optimal solution for the quadratic cost, we make use of the Kantorovich
duality relation and Brenier’s theorem, both of which are central and celebrated results in
the theory of optimal transportation. (A detailed and rigorous discussion of these results
can be found e.g. in the monograph in Ref. \[24\].) For the present situation, they relate the
formulation in Equation 1 to the following, “dual”, formulation of the optimisation problem,

\[
\hat{f}(y;\theta) = \arg \inf_{f \in \text{cvx}(Y)} \int dy \ q(y) f(y;\theta) + \int dx \ p(x|\theta) f^*(x;\theta).
\]  
(2)

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\[1\] One would generalise the line element of the Euclidean plane, \( ds^2 = dr^2 + r^2 d\phi^2 \), where \( r \) is the radial
coordinate and \( \phi \) the azimuthal angle. A weighted line element \( ds^2 = \alpha^2 dr^2 + r^2 d\phi^2 \) with \( \alpha > 1 \) would have
the desired effect: it defines a metric that is conformally equivalent to that of a cone with an opening angle
inversely proportional to \( \alpha \). Indeed, the generalisation of Equation 1 to Riemannian manifolds has been the
subject of intense study \[22\], and analogues of the results presented here for the case of the Euclidean plane
continue to hold. The practical solution of the optimal transport problem in this general setting introduces
additional technical challenges that are beyond the scope of this paper, but that can be addressed e.g. by
using the methods presented in Ref. \[23\].
The minimum in Equation 2 is taken over the set of functions $f(y; \theta)$ that are convex w.r.t. the argument $y$, which we denote as $f \in \text{cvx}(Y)$. The function $f^*(x; \theta)$ is the convex conjugate of $f(y; \theta)$ defined by

$$f^*(x; \theta) = \sup_{y \in Y} [x \cdot y - f(y; \theta)].$$

(3)

It is not immediately obvious that Equations 1 and 2 indeed produce identical solutions. We refer to Sections 1.1 and 2.1 (in particular Theorem 2.12) of Ref. [24] and also to Ref. [25] for a thorough explanation of all steps involved in a proof of their equivalence.

The optimal transport map $\hat{T}(x; \theta)$ that solves the original problem may be reconstructed from the function $\hat{f}(y; \theta)$ that attains the extremum in the dual problem in Equation 2,

$$\hat{T}(x; \theta) = \nabla x \hat{f}^*(x; \theta),$$

(4)

i.e. the transport function is a pure gradient field. The function $\nabla y \hat{f}(y; \theta)$ also has an interpretation as a transport function: it solves the “inverse” transport problem, i.e. maps $q(y)$ back onto $p(x|\theta)$ (see again Theorem 2.12 in Ref. [24]).

The optimisation in the original problem in Equation 1 is restricted to a very specific set of functions $T$, namely those for which $p'(x'|\theta) \equiv q(y)$. On the other hand, the dual formulation in Equation 2 is subject to the much more generic (and manageable) requirement that $f(y; \theta)$ be a convex function. This provides an avenue towards solving the dual problem numerically, and then reconstructing the optimal transport map by virtue of Equation 4.

4.3. A more practical form

To express Equation 2 in the form of a functional optimisation problem, we make use of a property of the convex conjugate. The argument $\hat{y}$ that maximises the right-hand side of Equation 3 is given by the gradient of the transformed function, i.e. $\hat{y} = \nabla_x f^*(x; \theta)$. We thus have $f^*(x; \theta) = x \cdot \nabla_x f^*(x; \theta) - f(\nabla_x f^*(x; \theta); \theta)$ and the convex conjugate $f^*$ can also be computed via

$$f^*(x; \theta) = \arg \sup_{g \in \text{cvx}(X)} x \cdot \nabla_x g(x; \theta) - f(\nabla_x g(x; \theta); \theta).$$

(5)

As noted in Ref. [25], the dual problem may then be rewritten as

$$\left(\hat{f}, \hat{g}\right) = \arg \inf_{f \in \text{cvx}(Y)} \sup_{g \in \text{cvx}(X)} \int dy \ q(y) \ f(y; \theta)$$

$$+ \int dx \ p(x|\theta) \ [x \cdot \nabla_x g(x; \theta) - f(\nabla_x g(x; \theta); \theta)].$$

(6)

This minimax problem can now be solved efficiently by standard optimisation methods. Since $\hat{g}(x; \theta) \equiv f^*(x; \theta)$, the optimal transport function is then given by

$$\hat{T}(x; \theta) = \nabla x \hat{g}(x; \theta).$$

(7)
4.4. Optimal transport for one-dimensional distributions

The results stated above remain true, of course, in the special case of one-dimensional probability distributions. However, the solution for this situation has very peculiar characteristics, and thus deserves a special mention.

In one dimension, gradients of convex functions—and thus the transport function $\hat{T}(x; \theta)$—are nondecreasing, i.e. the calibration never corrects the (now scalar-valued) events in different directions. Moreover, $\hat{T}(x; \theta)$ can be expressed explicitly as

$$\hat{T}(x; \theta) = Q^{-1} \circ P(x|\theta),$$

where $P(x|\theta)$ and $Q(y)$ are the cumulative distribution functions corresponding to the densities $p(x|\theta)$ and $q(y)$. A rigorous derivation of Equation 8 is given in Section 2.2 of Ref. [24]. This simply states that applying quantile matching (a very widely used technique) is equivalent to our (much more general) philosophy of calibration through optimal transport.

In the case where both distributions are Gaussians, i.e. $p(\cdot|\theta) = \text{No}(\mu_1(\theta), \sigma_1(\theta))$ and $q(\cdot) = \text{No}(\mu_2, \sigma_2)$, the transport function results in the well-known expression

$$\hat{T}(x; \theta) = x - \frac{\mu_1(\theta)}{\sigma_1(\theta)} \cdot \sigma_2 + \mu_2.$$  

5. Practical implementation of optimal transport

To solve the minimax problem in Equation 6, we replace the integrals by their respective batched estimates using the dataset $\mathcal{X}$, produced by the uncalibrated simulator, and the calibration dataset $\mathcal{Y}$. Choosing a suitable parametrisation for the functions $f(y; \theta)$ and $g(x; \theta)$ then turns the minimax problem into a tractable, finite-dimensional optimisation problem.

In this section, we focus on the case where all simulation parameters $\theta$ are continuous. They are typically associated with a prior distribution $p(\theta)$. The dataset $\mathcal{X}$ extracted from the simulator then consists of (weighted) samples of the form $(\{x_i, w_i, \theta_i\})_{i=1,...,N}$, obtained from the uncalibrated distribution $p(x|\theta)p(\theta)$. Clearly, the calibration dataset $\mathcal{Y}$ does not intrinsically depend on $\theta$, and so consists of samples $(\{y_j, \theta_j\})_{j=1,...,M}$ drawn from the product distribution $q(y)p(\theta)$.

We parametrise the functions $f(y; \theta)$ and $g(x; \theta)$ with two neural networks with parameters $\phi$ and $\psi$, respectively. The network architecture (explained below) is chosen such that the parametrisations $f_{\phi}(y; \theta)$ and $g_{\psi}(x; \theta)$ express manifestly convex functions w.r.t. the arguments $y$ and $x$ (as required by Equation 6), but that the dependency w.r.t. the simulation parameters $\theta$ can be arbitrary. This choice reflects the fact that, for continuous simulation parameters, the uncalibrated distribution $p(x|\theta)$—and also the functions $f(y; \theta)$ and $g(x; \theta)$—are expected to depend smoothly on $\theta$.

This gives rise to the loss functional in terms of the parameters $\phi$ and $\psi$,

$$L_{\text{trans.}}(\phi, \psi) = \frac{1}{M_b} \sum_{j=1}^{M_b} f_{\phi}(y_j; \theta_j) + \frac{1}{N_b} \sum_{i=1}^{N_b} w_i \left[ x_i \cdot \nabla_x g_{\psi}(x_i; \theta_i) - f_{\phi}(\nabla_x g_{\psi}(x_i; \theta_i); \theta_i) \right],$$

(10)
where $N_b$ and $M_b$ are the respective batch sizes. In terms of Equation 10, the optimisation problem becomes

$$
(\hat{\phi}, \hat{\psi}) = \arg \inf_{\phi} \sup_{\psi} \mathcal{L}_{\text{trans}}[\phi, \psi],
$$

and the optimal transport function is approximated by

$$
\hat{T}(x; \theta) \approx \nabla_x g_{\hat{\psi}}(x; \theta).
$$

5.1. Network architecture

In Ref. [26], a network architecture is constructed (referred to as “partially input convex neural networks”, or PICNNs) which meets the requirements set out above for the parametrisations of $f(y; \theta)$ and $g(x; \theta)$. The employed strategy consists in first finding an architecture that ensures convexity w.r.t. the function arguments, and then adding on top of it the most general structure to cover the dependency w.r.t. the parameters $\theta$.

The full architecture of a network implementing a generic function $h(z; \theta)$ (used to represent both $g(x; \theta)$ and $f(y; \theta)$) is visualised in Figure 3. It is defined through the recursion relations

$$
v_{i+1} = g_i^{(v)} \left( W_i^{(v)} \left[ v_i \odot \left( W_i^{(vu)} u_i + b_i^{(vu)} \right) \right] + W_i^{(z)} \left[ z \odot \left( W_i^{(zu)} u_i + b_i^{(zu)} \right) \right] + W_i^{(u)} u_i + b_i^{(u)} \right), \quad v_0 = 0, \quad h(z; \theta) = v_k,
$$

and

$$
u_{i+1} = g_i^{(u)} \left( W_i^{(u)} u_i + b_i^{(u)} \right), \quad u_0 = \theta.
$$
The key observation made in Ref. [26] is that a positive linear combination of convex functions remains convex, as does the composition of a convex function with a convex non-decreasing function. This can be realised with a setup akin to a conventional feedforward neural network with hidden nodes $v_i$, provided that the weights $W_i^{(v)}$ remain nonnegative at all times. No restrictions exist on the biases $b_i^{(v)}$. Additional “bypass” connections with weights $W_i^{(z)}$ route the inputs to each hidden layer, which increases the expressiveness of the network while maintaining its convexity properties.

The most general dependence on the parameters $\theta$ can be implemented by adding all terms that are not in conflict with the convexity requirements. This allows an auxiliary feedforward structure with $k-1$ layers. Its hidden layers $u_i$ can be coupled to the main branch by decorating them onto the $v_i$ (through weights $W_i^{(v)}$ and biases $b_i^{(v)}$) as well as the bypass connections (through $W_i^{(z)}$ and $b_i^{(z)}$). This is performed by means of the Hadamard (element-wise) product, denoted as $\odot$. To ensure that this does not compromise the non-negativity of the weights $W_i^{(v)}$, the decorated term itself must be nonnegative, indicated by the notation $(\cdot)_+$. The $u_i$ can also enter directly as an additive term with weights $W_i^{(u)}$.

As mentioned above, the activation functions $g_i^{(v)}(s)$ must be convex nondecreasing. No restrictions are posed on the activations of the auxiliary path, $g_i^{(u)}$. According to Equation 7, the transport function is determined by the gradient of $g_\psi$. To ensure a smooth transport, the $g_i^{(v)}$ should therefore also have a smooth first derivative. Using a parabolic function to smoothly connect the two line segments in a leaky ReLU gives rise to an activation function $\text{act}(s)$ which satisfies all requirements,

$$\text{act}(s) = \begin{cases} 
bs & s < 0 \\
bs + \frac{c-b}{2}s^2 & 0 \leq s \leq 1 \\
\frac{b-c}{2} + cs & 1 < s,
\end{cases}$$  \hspace{1cm} (13)$$

where $b$ and $c$ correspond to $\text{act}'(s)$ for $s < 0$ and $1 < s$, respectively. For this study we have chosen $b = 0.05$, $c = 1$, although these parameters were not optimised.

5.2. Training methodology

During training, the parameters of the networks expressing $f_\phi$ and $g_\psi$ are updated iteratively, similar to the joint training of the generator and adversary in Generative Adversarial Networks [27]. According to Equation 6, the training proceeds schematically as

$$\phi \leftarrow \phi - \eta_f \nabla_\phi \mathcal{L}, \quad \psi \leftarrow \psi + \eta_\theta \nabla_\psi \mathcal{L}.$$  \hspace{1cm} (14)$$

In practice, it can be beneficial to apply multiple update steps to $\psi$ for each $\phi$ update, and/or to use a larger learning rate, $\eta_\theta > \eta_f$.

To ensure convexity of $f(y; \theta)$ and $g(x; \theta)$, the nonnegativity of the corresponding weights $W_i^{(v)}$ must be maintained; to achieve this, a suitable regularisation term is added to the training loss, and the total loss function becomes

$$\mathcal{L}[\phi, \psi] = \mathcal{L}_{\text{trans.}}[\phi, \psi] + \sum_{W_i^{(v)} \in \phi, \psi} \max(-W_i^{(v)}, 0).$$  \hspace{1cm} (15)$$
Ensuring the nonnegativity of these weights in a “soft” manner by means of these regularisation terms was found to lead to a more stable training compared to forcing all negative weights to zero after each update step. This is in line with the observations made in Ref. [25].

6. Application in experimental particle physics

We explore three examples with increasing complexity in order to illustrate the effectiveness of our method in a situation inspired by experimental particle physics. While these examples are very simple and therefore easily visualised and understood, we emphasise that they are arbitrary and that the method described previously can be applied to similar problems of greater complexity and higher dimensionality.

6.1. Calibrating the signal prediction

In experimental particle physics, it is common to study the distributions of physically motivated event variables. One such variable is the invariant mass of a system of particles, for which the physical process of interest, the “signal”, typically produces a distinct peak in the observed data, corresponding to a short-lived particle created in the collision.

Here we consider a signal process characterised by a density that is the sum of two normal distributions: a “core” with \( \mu = 0.1 \) and \( \sigma = 0.1 \) and a “tail” with identical mean, but \( \sigma = 0.25 \). In our example, we suppose the core and the tail distributions come in equal proportions. This corresponds to the common scenario in which there are several detector effects that determine the overall resolution on the observed mass variable \( y \).

We assume that these resolution effects are not well-known \textit{a priori} and that they are thus not modelled by the simulation: it predicts the uncalibrated mass variable \( x \) to follow a single normal distribution with \( \mu = -0.5 \) and \( \sigma = 0.25 \). It is then the objective of the transport function \( x' = \hat{T}(x) \) to rectify the discrepancy in the shape of the peak between the simulator prediction and the observed data. For this simple example, we take the simulation to have no additional parameters, i.e. the argument \( \theta \) is absent.

To construct the transport function, we follow the procedure described in Section 5. The architecture of the networks parametrising \( f_\phi(y) \) and \( g_\psi(x) \) is that of Figure 3. We use relatively small networks, with three hidden layers of 32 \( v_i \) nodes each, using the PyTorch Deep Learning Library [28]. The \( u_i \) nodes, corresponding to non-convex inputs, are not required in this case.

Figure 4 shows the transport function resulting from following this procedure, as well as a comparison of the data and simulated distributions before and after the calibration is applied. Excellent closure is obtained after about 30 minutes of training on a laptop computer using batches of \( N_b = M_b = 256 \) samples, the Adam optimiser [29], and learning rates of \( \eta_f = \eta_g = 10^{-4} \); in this case we update the parameters \( \psi \) ten times for each \( \phi \) update. These hyperparameters were chosen by performing a coarse scan of batch size (from 16 to 1024), learning rate (from \( 10^{-3} \) to \( 10^{-5} \)), and number of nodes per layer (from 8 to 64), and choosing the parameters with the smallest binned \( KL \)-divergence between transported simulation and calibration data.
Figure 4: The functions (a) $f_\phi(y)$, (b) $g_\psi(x)$, and (c) the resulting transport $\hat{T}(x) = \nabla_x g_\psi(x)$ after the training procedure for a signal-only scenario. The original simulated prediction, $p(x)$, and transported prediction, $p'(x')$, are compared to the calibration data distribution, $q(y)$, in (d), with closure observed at the 5% level in each bin after the calibration is applied.
6.2. Calibration in the presence of background processes

It is not typically the case that a sample of events is available which is perfectly pure in the signal process. Instead, contributions from *backgrounds* are also important, i.e. physical processes that are not immediately of interest, but also contribute to the collection of studied events, and that must thus be understood. In many cases, background processes lead to distributions over the observed mass variable $y$ that are rather broad and smooth, and do not contain any sharp features.

To illustrate how backgrounds can be accommodated in our formalism, we now suppose that the observed data are actually drawn from two distributions: the signal events still come from the double-Gaussian described in the example in Section 6.1, but 50% of all events originate from a background process for which $y$ follows a normal distribution with $\mu = -0.5$ and $\sigma = 1$.

We suppose that the simulation of the backgrounds is already accurate, and does not require calibration. The objective then becomes to construct a transport function $\hat{T}(x)$ that ensures that the simulation of the signal is properly calibrated.

This means that we require $\hat{T}(x) = x$, whenever $x$ corresponds to a simulated background event. To achieve this, we define

$$
\hat{g}_\psi(x) = \begin{cases} 
    g_{\psi}^{\text{sig}}(x) & \text{for signal events} \\
    \frac{1}{2}x^2 & \text{for background events},
\end{cases}
$$

noting that $\nabla_x \frac{1}{2}x^2 = x$ for background events, as required. The function $g_{\psi}^{\text{sig}}(x)$ remains unspecified and is adjusted during the training procedure in the usual way to yield a minimal modification of the signal simulation that results in its calibration.

Equation 16 goes a rather direct route towards constructing a transport function with the desired properties. Appendix A shows that this form of $\hat{g}_\psi(x)$ is also formally correct, i.e. it effectively leads to a subtraction of the background prediction from the total distribution contained in the calibration dataset in order to isolate and calibrate the signal.

We proceed with the minimisation over the network parameters $\phi$ and $\psi$ in the usual way to obtain $f_\phi$ and $g_\psi$. Also the network architecture introduced in Section 6.1 remains unchanged. Figure 5 shows the signal transport function $\nabla_x g_\psi^{\text{sig}}(x)$ and comparisons between data and simulation before and after the calibration is applied. Excellent closure is again observed between observed data and the calibrated simulator prediction.

6.3. Systematic uncertainties

We now introduce additional simulation parameters $\theta$, which control the exact shape of the simulated signal distribution. By virtue of their prior distribution $p(\theta)$, the simulation parameters give rise to a notion of systematic uncertainties on the prediction of the uncalibrated simulation. In this context, these parameters are often referred to as “nuisance parameters” since they are not of particular (physics) interest, but they nevertheless affect the simulation, and hence its calibration.

To illustrate the handling of systematic uncertainties in our calibration strategy, we return again to our example from Section 6.2. We extend it by introducing a single new
Figure 5: (a) the transport function $\hat{T}(x) = \nabla_x g_{\hat{\psi}}(x)$ after the training procedure for a signal-plus-background scenario. The original simulated prediction, $p(x)$, and transported prediction, $p'(x')$, are compared to the calibration data distribution, $q(y)$, in (d), with closure observed at the 5% level in each bin after the calibration is applied.
parameter, $\theta_0$, which alters the simulator such that the predicted centroid and width of the signal distribution depend linearly on it:

\[
\mu(\theta_0) = -0.5 + 0.1 \theta_0, \\
\sigma(\theta_0) = 0.25 + 0.1 \theta_0.
\]

The functions $f_\phi(y; \theta)$ and $g_\psi(x; \theta)$ now make use of the full neural network architecture shown in Section 5.1. We assume the prior distribution $p(\theta_0)$ to correspond to a standard normal distribution, as is common in experimental analyses [30]. This imposes a Gaussian constraint on $\theta_0$, with $\theta_0 = 0$ corresponding to the “nominal” best-guess value of the parameter. The training proceeds according to the loss function in Equation 10, and during training we sample values of $\theta_0$ according to a uniform distribution in the range $(-2, 2)$. With the Gaussian constraint, this ensures good behaviour at the $2\sigma$ level; if stability is needed beyond $2\sigma$, this is achievable by extending the training range of $\theta_0$.

Due to the added complexity needed to obtain closure for a range of $\theta_0$, we use slightly more complex networks to implement $f_\phi$ and $g_\psi$ in this case. Using the nomenclature from Figure 3, the networks are composed of three hidden layers, with 64 $v_i$ nodes and 16 $u_i$ nodes in each layer.

The simulator and data distributions as well as the corresponding transport functions and their dependence on $\theta_0$ are shown in Figure 6. Good agreement is achieved between the target dataset and transported simulation for a wide range of $\theta_0$ values, and the resulting $\hat{T}(x; \theta_0)$ defines a family of transport calibrations that can be used to evaluate the uncertainty on the calibration $\hat{T}$ when applied in a signal region of interest.

7. Conclusions and outlook

The use of calibration data to correct details of stochastic simulators is an integral component of the experimental physicist’s toolkit. We argue that the language of transportation theory is well-suited for describing the adjustment of predictive models based on experimental control samples. Calibrations derived using optimal transport methods are guaranteed to minimally alter simulated predictions, which contain years of collected domain knowledge, while faithfully reproducing relevant distributions observed in calibration data.

We have presented a concrete implementation plan for obtaining calibration transport maps, which are realised via feed-forward neural networks that are straightforward to train on modern computing hardware. Illustrative examples were shown, demonstrating the capability of the method and its ability to handle background processes and systematic uncertainties on the simulated prediction. Calibrations derived as transport maps and implemented as neural networks overcome a number of shortcomings inherent to current methods, which are less general, do not scale well into high-dimensions, or do not preserve some important invariants such as the total number of predicted events. Based on the generality of this approach and the fact that it depends on an already well-established technology (neural networks), our method promises wide applicability across experimental physics.
Figure 6: (a) the transport function $T(x; \theta_0) = \nabla_x g_\psi (x; \theta_0)$ after the training procedure for a signal-plus-background scenario. The original simulated prediction, $p(x)$, and transported prediction, $p'(x')$, are compared to the target data distribution, $q(y)$, in (b), with closure at the 10% level observed in each bin after the calibration is applied for each of the three $\theta_0$ points of interest.
Optimal transport also provides a rigorous framework in which to generalise the calibration method presented here. In general, one can allow the simulator to produce data that populate a Riemannian manifold \[22, 23\], on which a natural cost function is given by the metric distance integrated along geodesics. Taking the global structure of the data into account is important for the proper treatment of components that are compact or periodic, e.g. angular directions. Furthermore, it can be useful to decouple the distribution that is being calibrated (e.g. the di-jet invariant mass) from the variable to which the calibration is applied (e.g. the transverse momenta of the jets). In effect, this defines a relaxation of the original Monge problem in Equation 1. It applies very naturally to situations where corrections should be applied directly to the primary output of the simulator, but only the proper calibration of derived quantities (possibly of lower dimension) is important.

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A. Calibrations in the presence of backgrounds: two equivalent ways

We consider the calibration of the signal in the presence of backgrounds. The latter are assumed to be modelled correctly by the simulation, i.e. do not require a calibration. The distribution delivered by the uncalibrated simulator is thus

\[ p(x; \theta) = (1 - \mu) p_{\text{bkg}}(x) + \mu p_{\text{sig}}(x; \theta), \]  

(17)

while the target distribution for the calibration is

\[ q(y) = (1 - \mu) p_{\text{bkg}}(y) + \mu q_{\text{sig}}(y). \]  

(18)

The coefficient \( \mu \) corresponds to the signal fraction.

A.1. One solution

We thus want to construct a transport function that corrects the simulation of the signal, so that

\[ p'(x' \mid \theta) \equiv q(y). \]  

According to Equation 6, the corresponding transport function is the solution of the minimax problem

\[ \begin{align*}  
(\hat{f}, \hat{g}) = \arg \inf_f \sup_g \int dy \ q_{\text{sig}}(y) f(y; \theta) + \int dx \ p_{\text{sig}}(x \mid \theta) \left[ x \cdot \nabla_x g(x; \theta) - f(\nabla_x g(x; \theta); \theta) \right],  
\end{align*} \]  

(19)

where, as usual, \( f \in \text{cvx}(Y) \) and \( g \in \text{cvx}(X) \). To make the following manipulations more concise, we use the shortcut \( \mathbb{E}_h[k(x)] \) to denote the expectation of \( k(x) \) under the density \( h(x) \), i.e. \( \mathbb{E}_h[k(x)] = \int dx \ h(x) k(x) \).

To express Equation 19 in terms of the densities \( q_{\text{sig}}(y), p_{\text{bkg}}(y) \) and \( p_{\text{sig}}(x; \theta) \), we make use of the relation

\[ \mathbb{E}_{q_{\text{sig}}} = \frac{1}{\mu} \mathbb{E}_q - \frac{1-\mu}{\mu} \mathbb{E}_{p_{\text{bkg}}}, \]

which follows from the identity

\[ q_{\text{sig}}(y) = \frac{1}{\mu} q(y) - \frac{1-\mu}{\mu} p_{\text{bkg}}(y). \]

With this, we find

\[ (\hat{f}, \hat{g}) = \arg \inf_f \sup_g \frac{1}{\mu} \mathbb{E}_q [f(y; \theta)] - \frac{1-\mu}{\mu} \mathbb{E}_{p_{\text{bkg}}} [f(y; \theta)] + \mathbb{E}_{p_{\text{sig}}} [x \cdot \nabla_x g(x; \theta) - f(\nabla_x g(x; \theta); \theta)], \]

A.2. The same solution again

Alternatively, we can start from the original transport between \( p(x; \theta) \) and \( q(y) \),

\[ (\hat{f}, \hat{g}) = \arg \inf_f \sup_g \mathbb{E}_q [f(y; \theta)] + \mathbb{E}_p [x \cdot \nabla_x g(x; \theta) - f(\nabla_x g(x; \theta); \theta)] \]  

(20)

19
and insert the parametrisation from Equation 16 for $g(x; \theta)$, i.e. demand that $g(x; \theta) = \frac{x^2}{2}$ whenever it is evaluated on a background event.

The above minimax problem then turns into

$$(\hat{f}, \hat{g}) = \arg \inf_f \sup_g E_q[f(y; \theta)] + (1 - \mu)E_{p_{bkg}}[y \cdot y - f(y; \theta)]$$

$$+ \mu E_{p_{sig}}[x \cdot \nabla_x g(x; \theta) - f(\nabla_x g(x; \theta); \theta)].$$

Dropping the term $(1 - \mu)E_{p_{bkg}}[y \cdot y]$, which is independent of $f$ and $g$, we get,

$$(\hat{f}, \hat{g}) = \arg \inf_f \sup_g E_q[f(y; \theta)] - (1 - \mu)E_{p_{bkg}}[f(y; \theta)] + \mu E_{p_{sig}}[x \cdot \nabla_x g(x; \theta) - f(\nabla_x g(x; \theta); \theta)],$$

which is identical to the result from Section A.1.
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