In this study we develop a 32-year (1982–2013) otolith biochronology of the commercially important deepwater hake *Merluccius paradoxus* in the northern Benguela, Namibia. Mean annual growth (mm) calculated from 140 thin-sectioned *M. paradoxus* otoliths were compared with change in mean length at age 3 to age 4 determined from Namibian whole-otolith-read age-length keys (ALKs). Annual growth rates calculated from the two methods (overlapping 2000–2013) were strongly positively correlated ($\rho = 0.730$, $n = 14$, $p < 0.01$). This indirectly validated annual age determination of *M. paradoxus*, the accuracy of otolith chronologies, and the ability of ALKs to capture annual variability in fish growth. Annual *M. paradoxus* growth rates were significantly positively correlated with the July–September upwelling index 1982–2013 at 30$^\circ$S, ($\rho = 0.414$, $n = 32$, $p < 0.05$) and positively correlated with August mean chlorophyll-a concentrations (as indicator for primary production) 2002–2013 in the 28–30$^\circ$S area ($\rho = 0.734$, $n = 12$, $p < 0.01$). Annual *M. paradoxus* growth rates significantly negatively correlated with October (austral spring) sea surface temperatures in the 24–28$^\circ$S area ($\rho = -0.381$, $n = 32$, $p < 0.05$). This Orange-River Namqua upwelling cell corresponds to the area where juvenile and young adult *M. paradoxus* live, suggesting growth rate strongly responds to local forcing. We also determined that mean length-at-age 3 calculated from ALKs (current and literature) significantly increased from 1977 through 2016 at 0.075 cm.year$^{-1}$ ($t = 3.04$, df = 41, $p = 0.004$), while length-at-age 8 significantly decreased at 0.25 cm.year$^{-1}$ ($t = -3.59$, df = 30, $p = 0.001$). Both trends may indicate fisheries-induced adaptive changes. *M. paradoxus* occurring at >300 m bottom depth, are thus strongly influenced by fisheries. As an upper-level demersal predator, this species integrates signals throughout...
Merluccius paradoxus (Hutchings et al., 2009; Verheye et al., 2016). The deepwater environment typified by a complex ecosystem structure that supports numerous species of substantial commercial value in the west coast of southern Africa, supports a productive Benguela Upwelling System, located along the southern Benguela characterized by strongly seasonal upwelling (Hardman-Mountford et al., 2003; Lamont et al., 2018a). Recent studies have revealed that upwelling ecosystems globally are undergoing long-term change, marked by a poleward shift in upwelling, linked with poleward shifts in the atmospheric high pressure systems that drive upwelling-favorable winds (Rybak, 2015; Wang et al., 2015). Similar changes have been observed in the Benguela system, with substantially less upwelling-favorable wind in the northern Benguela in recent years (Lamont et al., 2018a), accompanied by positive trends in Sea Surface Temperatures (SSTs) (Jarre et al., 2015). In contrast, coastal cooling of about 0.5°C per decade has been identified in the southern Benguela, though this cooling has been less pronounced farther south on the Agulhas Bank (Rouault et al., 2010). These cooling trends appear to be linked to an increase in upwelling-favorable south-easterly and easterly winds (Rouault et al., 2010; Lamont et al., 2018a).

Water temperature is one of the main factors influencing the biology, distribution and life history of fishes, along with other aquatic organisms, especially in high latitude freshwater systems (e.g., Magnuson et al., 1997; Casselman et al., 2002; Vieira et al., 2020). In temperate areas, too, fish growth increased due to a disproportionate lengthening of the growing season with climate change (e.g., Morrongiello et al., 2011). In the northern part of the northern Benguela (Angola-Benguela Front), a temperate marine system, increasing inshore temperatures in an ocean-warming hot-spot resulted in species mixing and even hybridization of two Sciaenid Argyrosomus species (Potts et al., 2014). Changing water temperatures thus have consequences for processes such as growth rates, timing of seasonal events (phenology), species distributions, species interactions, and overall levels of productivity (e.g., Carey and Zimmerman, 2014; Morrongiello et al., 2014; Tanner et al., 2019).

Apart from changes in the life history strategies and productivity of fish populations due to climate, changes in life history strategies can also occur as a result of selective fishing. For example, Morrongiello et al. (2019) determined, using a biochronology approach on a recently exploited temperate marine reef fish in Tasmania, that fishing not only increased adult growth rates, but it also changed thermal reaction norms to increased sensitivity of these fish to temperature, over the relatively short time period of 10 years. Fisheries-induced changes in growth rates due to density-dependent processes or selection for investment into pre-mature growth and reproduction of fish, were discussed by Lorenzen and Enberg (2002), and Enberg et al. (2012), respectively.

In the northern Benguela, the hakes, consisting of the shallow-water hake Merluccius capensis and deepwater hake *M. paradoxus* make up Namibia’s most valuable fishery, and are a research priority at the Ministry of Fisheries and Marine Resources (MFMR). Exploitation of hakes (demersal trawl-fishing) in Namibia started in 1964 with catches peaking at 800,000 tons in 1972, with open access fishing by many foreign fleets. This resulted in an initial drastic decline in the stock biomass (Kirchner et al., 2012). From 1976 to 1989 the fishery was controlled by the International Commission for Southeast Atlantic Fisheries (ICSEAF), implementing a legal mesh size limit and member country quotas or total allowable catch (TAC) limits. Despite this, the stock biomass continued to decline to only 170,000 t by 1980. In 1990 Namibia became independent and the Ministry of Fisheries and Marine Resources (MFMR) took over management of the fishery on a heavily depleted resource. More stringent management measures were immediately implemented including closure of fishing to all foreign vessels, enforcement of a 200-mile exclusive economic zone and immediately reducing the TAC to 60,000 tons in 1990-1991. From 1992 to current, between 88,000 and 190,000 t (mean of 148,000 t) of hake have been caught annually (Wilhelm et al., 2015a). The Namibian hake resource is currently assessed (two stocks combined) using a Statistical Catch-at-AGE Analysis (SCAA, since 1998) at between 14 and 26% of the pristine spawning stock biomass (SSB) and at 95–120% at best case scenario of SSB values of 1990 (Kirchner et al., 2012; Kathena et al., 2016). Climate-induced as well as fisheries-induced changes in growth, are expected to be evident in these stocks over the last few decades.

*Merluccius paradoxus* are believed to spawn on the western Agulhas Bank between 32 and 34°S in the southern Benguela (Strømme et al., 2016). The smallest juveniles caught in bottom trawls (<5 cm total length), are usually found between 31 and 32°S off the West coast of South Africa. As these fish grow, they move northwards and southwards, and fish between 15 and 65 cm...
total length are most frequently caught off Namibia (Strømme et al., 2016). *M. paradoxus* grow to about 100 cm in length and to about 12 years of age. Length at 50% maturity (L50) is 41–54 cm and age at 50% maturity (A50) 4–6 years (Durholtz et al., 2015; Wilhelm et al., 2019).

Growth rates and age information of fishes are imperative to understanding population dynamics and life history strategies, and ultimately to improve methods for sustainable exploitation. Such information can be obtained through analyzing annual growth increments in fish otoliths. There are two ways to use annual otolith increments to describe fish growth rates. Counting annual increments to determine the age of individual fish is a relatively simple method to calculate population-level growth rates using, e.g., von Bertalanffy growth parameters. This is mostly applied to incorporate individual or cohort-level growth rate effects (e.g., Pilling et al., 2002), but also to explore environmental and temporal trends using time-series data (e.g., Baudron et al., 2014). This approach is usually used at a population-level and time-invariant analysis, and misses the complexity of the drivers of growth at individual level (Morrongiello and Thresher, 2015).

The second approach is using annual otolith growth as a proxy of fish growth and relies on the fact that fish somatic growth (fish length) and otolith growth (otolith length) are strongly related. The width of each annual increment can be measured to estimate the annual growth of an individual fish over their life span. An annual growth increment on a fish otolith usually consists of one zone pair, one translucent and one opaque zone. The translucent zone is thought to form during a period of slow somatic growth and the opaque zone is thought to form during a period of fast somatic growth, which often coincides with the productive period. The width of the annual increment thus generally reflects the annual fast growth period of the fish. Consequently, otolith chronologies have been used to describe annual growth rate variability of fish and explore its environmental drivers. For example, Boehlert et al. (1989) and Dorn (1992) showed that increment width chronologies of Pacific whiting, splinitose rockfish and canary rockfish caught off the coast of California were significantly correlated to sea surface temperature and upwelling, amongst other environmental variables. In this study we used both approaches to describe changes in *M. paradoxus* growth over time.

Black et al. (2005) described how the dendrochronology (tree ring analysis) technique of crossdating could be applied to sclerochronology (analysis of growth increments in hard structures) using a long-living rockfish species caught off the Californian coast. The technique assumes that climate limits growth variation over time, causing synchronous growth increment widths of individuals from the same species and climate regime. Crossdating is then the process of matching these synchronous growth patterns among individuals beginning at the known year of capture and working back through time. Resulting biochronologies are therefore annually resolved and can be readily integrated with climate data. In the California Current Ecosystem of the northeast Pacific, Black et al. (2011) demonstrated that Pacific rockfish otolith chronologies strongly reflect inter-annual variability in winter upwelling. Moreover, these rockfish chronologies strongly correlated with other biological indicators including time series of seabird reproductive success, corroborating their ability to capture climate variables associated with bottom-up forcing (Black et al., 2014). Whenever possible, all growth increment data should be crossdated prior to developing biochronologies.

Morrongiello et al. (2014) and Morrongiello and Thresher (2015) use mixed effect modeling (MEM) (Weisberg et al., 2010) to analyze growth-increment widths. MEM efficiently represents the hierarchy of otolith biochronology data (Morrongiello et al., 2014) and captures both extrinsic (time, temperature, fishery activity, and spatial structure) and intrinsic (individual, age, sex, and cohort) effects. The MEM method makes use of all available information to partition the relative impact of intrinsic and extrinsic effects on growth variability (Morrongiello and Thresher, 2015). Morrongiello and Thresher (2015) demonstrate that temporal growth rate variability in Australian tiger flathead *Platyccephalus richardsoni* was driven by two main factors, namely annual fluctuations in environmental conditions (extrinsic) and an intrinsic cohort-specific factor that reflects density dependence and/or juvenile experience. The biochronology approach thus demonstrates important applications in broad-scale ecological connectivity and hind-casting (Black et al., 2011, 2014) as well as disentangling fisheries, environmental, and genetic effects (Morrongiello et al., 2014; Morrongiello and Thresher, 2015) and in this way enables forecasting expected responses, which is necessary for Ecosystem Approach to Fisheries (EAF) management strategies. Otolith archives and their associated data also exist for the main commercial fish species in Namibia since 1991.

This study thus aims to make use of these archived otolith resources of the heavily exploited deepwater hake, *M. paradoxus*, in Namibia. Of the two hake species in the Benguela *M. capensis* has relatively fast growth rates and considerable dating uncertainties that result in age over-estimates when enumerating otolith translucent zones (Wilhelm et al., 2013, 2015b, 2017, 2018, 2019). We therefore used *M. paradoxus*, given that otolith growth and zonation is less variable, likely due to the stability of their deeper habitat relative to *M. capensis* (Wilhelm et al., 2019). Accordingly, ages from their otoliths are generally more readily identifiable due to fewer “false zones” or “checks” on their otoliths (BCC, 2015).

The aim of this study was, first, to develop an otolith biochronology of the commercially important deepwater hake, *M. paradoxus* caught in the northern Benguela off Namibia using otoliths collected from 1991 to 2014. We aimed to subsequently correlate the predicted growth estimated from the biochronology with years 3–4 growth rate changes of the fish from age-length keys, and fish condition indices to validate *M. paradoxus* annual age determination. We additionally aimed to test the relationships between predicted annual growth form the biochronology and selected environmental/climatic factors. Lastly, we aimed to describe the long-term changes in length-at-age of *M. paradoxus* using the first approach to age determination with current and historical data (from which otoliths had not been archived).
MATERIALS AND METHODS

The Study Species and Sample Collection

*Merluccius paradoxus* were collected during routine biomass surveys conducted by the Ministry of Fisheries and Marine Resources, Namibia (MFMR) covering the Namibian coast from 17 to 29°S (Figure 1 and Table 1). Otoliths were available from surveys conducted since 1991 and ±30 otoliths were randomly selected from fish ≥35 cm, within ±5-year intervals from 1991 through 2013, totalling an initial 258 otoliths. Because of the truncated stock structure and distribution of *M. paradoxus* as described by Stromme et al. (2016) not many >65 cm *M. paradoxus* were available from survey catches in Namibia (Table 1).

*Merluccius paradoxus* were assumed to spawn in austral spring to summer (September to February) (Jansen et al., 2015). Thus the translucent zones on the otoliths of *M. paradoxus* were assumed to form mainly in spring (September–October), which has been validated with edge analysis (ICSEAF, 1983a; BCC, 2015; MW, unpublished data). In addition, most of the fish were caught January-February (Table 1), and most of them had translucent zones near or on the edge of the otolith, indicating that they had just completed these in spring or early summer or were still completing them in late summer. One increment was designated as the couplet of the relatively wide opaque zone and the relatively narrow translucent zone. Measurement of the last increment, if complete, was thus assumed to be from the austral spring-early summer (near the end of the previous year) to the summer of the year before, and thus reflected growth of mostly that calendar year (over austral autumn, winter and spring). For example, if a fish was caught in February 2014, the last complete increment was assumed to reflect growth from October 2012 to October 2013, and the second-last from October 2011 to October 2012, continuing backward in time. If incomplete, the last increment was omitted from the analysis.

Otolith Chronologies

In the laboratory, the otoliths were embedded in epoxy resin and transversally sectioned through the nucleus at 0.5 mm thickness using a low-speed double-bladed saw. The sections were mounted onto glass slides using Crystalbond™ adhesive. After drying, otoliths were polished with 10 µm lapping film to improve the clarity of the macrostructures. Each otolith was digitally photographed using Image-Pro® Plus Version 6.0 (Media Cybernetics, Inc.) software under a Leica dissecting microscope at 3.2× magnification.

Once imaged, the width of each annual otolith growth increment was measured using ImageJ 1.49d¹ (Abramoff et al., 2004) and the macro ObjectJ plugin Version 1.03p.² ImageJ was used here because of free access and accessibility for multiple

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¹http://imagej.nih.gov/ij/
²https://sils.fnwi.uva.nl/bcb/objectj/examples/TreeRings/TreeRings-9.htm
TABLE 1 | Sample sizes (N), age and length ranges of Merluccius paradoxus used for the final dataset of the otolith biochronology, sampled during Namibian hake biomass surveys 1991 to 2014.

| Year | Months | N | Latitude range (°S) | Fish length range (cm) | Fish age range (years) |
|------|--------|---|---------------------|-----------------------|------------------------|
| 1991 | 10–11  | 15 | 20–28               | 35–55                 | 3–6                    |
| 1992 | 10–11  | 8  | 21–27               | 38–59                 | 3–12                   |
| 1994 | 10–11  | 2  | 23, 25              | 43, 78                | 5, 9                   |
| 1997 | 1–2    | 31 | 18–29              | 35–84                 | 4–15                   |
| 2002 | 1–2    | 27 | 18–27              | 35–70                 | 4–12                   |
| 2008 | 1–2    | 23 | 23–29              | 35–65                 | 3–10                   |
| 2013 | 1–2    | 12 | 26–29              | 50–82                 | 6–16                   |
| 2014 | 1–2    | 22 | 25–29              | 36–70                 | 4–11                   |

Mixed Effects Modeling
Once increments were marked and cross-dated, each fish was assigned a final age-at-capture (AAC) and Cohort. Cohort was defined as the year of birth of the fish (calculated for February-caught fish as: year of capture – AAC – 1). Each raw otolith increment width measurement in mm (Inc) was assigned an age at the time of increment formation (Age) and calendar year of increment formation (Year). Before analysis, otolith increment-widths, Age, and AAC were log-transformed and mean-centered. These data were then used in a mixed effects model, to relate increment width (Inc), the response variable, to a series of intrinsic (fish-specific) and extrinsic (environmental, represented by inter-annual variation) factors (Morrongiello and Thresher, 2015). Intrinsic factors included Age, AAC, Sex (of the individual fish, male or female), allowing for the interaction between Age and Sex, measurement Axis (“Long” or “Short,” see Figure 2) on the otolith, allowing for the interaction between Axis and Age, Cohort and FishID (each individual sample’s unique identification). Age, AAC, Sex and Axis were treated as fixed effects. The effects of Cohort and FishID intercepts were treated as random effects. FishID accounts for pseudo-replication, i.e., correlation of the same individual due to genetic ID and sample preparation. Cohort accounts for density dependence and/or juvenile experience. The AAC term was included to correct for biases associated with sampling (skewed age distributions) and tests for bias associated with growth-rate based selectivity (Morrongiello and Thresher, 2015). Year intercept was treated as a random effect. Further random effects included were by-FishID, by-Cohort, and by-Year random slopes for the effects of age. Equation 1 shows the full model fitted to the entire otolith chronology data series 1982–2012 (Table 1).

\[
\log(\text{Inc}) \sim c.(\log(\text{Age})) + \text{Age} \times \text{Sex} + c.(\log(\text{AAC})) + \text{Age} \times \text{Axis} + c.(\log(\text{Age}))|\text{FishID} + c.(\log(\text{Age}))|\text{Cohort} + c.(\log(\text{Age}))|\text{Year} \tag{1}
\]

A series of mixed effects models (Weisberg et al., 2010; Morrongiello and Thresher, 2015) was used to test for the best possible combination of the explanatory variables, testing all possible combinations of fixed effects and random effects, intercept and slope for random effects [denoted by e.g., \(\log(\text{Age})|\text{FishID}\), Equation 1]. The most parsimonious model was selected using maximum likelihood estimates of error and the lowest Akaike’s Information Criterion (AIC) corrected for small sample sizes (AICc) (Burnham and Anderson, 2002; Zuur et al., 2009). The optimal selected model was refitted with restricted maximum likelihood (REML) to produce unbiased parameter estimates (Zuur et al., 2009), using the random effects of year (intercept) to produce a best linear unbiased predictor of growth (BLUP) for each year 1982–2012. The BLUP time series was used to visualize the temporal pattern of \(M. \ paradoxus\) growth variability and was correlated with an independently calculated growth rate, a fish condition index as well as environmental data/indices that represent local and remote forcing. All analyses were performed in R Version 3.5.0 (R Core Team, 2018), with libraries lme4 (Bates et al., 2015), AICcmodavg (Mazerolle, 2019), effects (Fox, 2003; Fox and Weisberg, 2019), and lattice (Sarkar, 2008) for plotting.

Annual Population-Level Age 3 Growth
Annual age-length-keys (ALKs), used in the Namibian hake stock assessments and read from whole otoliths collected in January-February of each year, were available from 2000 to 2016 (SP, MFMR, unpublished data). These were used to calculate mean lengths at each age 1 to 7 years for each calendar year (Supplementary Table S1). These data were used to calculate change in length from age 3 to 4 years (G3) from calendar year \(y\) to the next, at the population level, by the following equation:

\[
G3 = L4_{y+1} - L3_y \tag{2}
\]

where \(y\) = the year of growth, \(L3_y\) = length at age 3 in calendar year \(y\), and \(L4_{y+1}\) = length at age 4 in calendar year \(y + 1\) (Supplementary Table S1). Growth from age 3 to 4 years was used because these fish are in the juvenile-to-sexually mature adult transition (Wilhelm et al., 2015a) and are no longer influenced by cohort effects or hatch timing (this study, data not shown).
**Fish Condition Index**

Biological data collected from Namibian hake biomass surveys from 1994 to 2013 (Table 2) were used to calculate mean relative weight as a measure of fish condition for that year for M. paradoxus. Relative weight was calculated by dividing by survey-measured wet weight of each individual fish by the expected weight. Expected weight was calculated using a length-weight relationship:

$$\log(W_e) = b \times \log(L_t) + a$$  \hspace{1cm} (3)

where $W_e$ = expected weight in g, $L_t$ = total fish length in cm, and $a$ and $b$ are constants calculated for all fish measured for all surveys 1999–2013, $a = 3.10$ and $b = -2.31$ (Supplementary Figure S1).

After the data were cleaned for obvious errors using weight-length plots, the mean relative weight was calculated for each survey 1994 to 2013 for all fish between 25 and 59 cm. Most surveys were conducted in Quarter 1, January and February, while before 1997 surveys were also conducted in Quarter 2, April and May and Quarter 4, September to November (Table 2). Prior to 1997 not all years had surveys in the first quarter of the year, and there may also have been area and length sampling differences between these surveys. Thus, a General Linear Mixed Effects Model (GLMM) was performed in order to calculate a mean relative weight (condition index) for each year (keeping year as a random effect), taking into account all other variables (the unbalanced sampling design).

The full model, which was also the best-fit model (AICc = -78582.9, Supplementary Table S2) was as follows:

$$\log(W) \sim c.(\log(L_t)) \times \text{Sex} + 1/\text{Year} + \text{Quarter} + \text{Lat}$$  \hspace{1cm} (4)

where $W$ = wet weight of fish $i$ in g, $L_t$ = total length of fish $i$ in cm (continuous variable). All other factors were added as categorical variables: Sex = male or female or unsexed, Lat = latitude to the nearest degree in which fish $i$ was caught (17°–29°S), Year = year of sampling (1994–2013), Quarter = sampling quarter of the year, either 1 (summer), 2 (autumn), or 4 (spring) using the same centring function, model selection procedures and R-packages as explained for Equation 1. Within the selection procedure, Quarter and Lat were also added as random effects to test for
the best fit model. The model-estimated random intercept for each year was antilogged to obtain a relative weight for that year to be used for comparison with the otolith chronology annual growth estimate.

Environmental Data
The BLUP time series from the mixed effect model was correlated (Pearson’s Correlation Coefficient) with an upwelling index, chlorophyll-a concentration, and sea surface temperature (SST) as indicators of local environmental drivers. The environmental variables were used as monthly, three-monthly and annual means and were correlated with the annual BLUP they overlapped with (lag 0 and lag 1 for spring), as well as lagged by 1 year for all (lag 1 and lag 2 for spring).

Similar to Lamont et al. (2018a), daily NCEP-DOE Reanalysis 2 wind vectors (Kanamitsu et al., 2002) along the coast were used to calculate daily values of Ekman transport (m3 s−1) per 100 m of coastline at the Lüderitz upwelling cell (27°30′S) in the northern Benguela (15–29°S), and at the Namaqua upwelling cell in the southern Benguela (30°S). Daily values of cumulative offshore transport were summed per month to obtain total monthly values, which were then used as an indicator of upwelling variability.

Chlorophyll-a (chl-a) concentrations were calculated in mg m−3 using MODIS satellite data from the Giovanni website1 for the period 2002–2013 at 4 km resolution (Acker and Leptoukh, 2007). The chl-a data were summarized to calculate monthly means for each of six zones within the northern part of the Benguela ecosystem. These were defined by homogeneity in severity and seasonality of the Chl-a concentration as: Zone 1: the Angola–Benguela Front (14–19°S, 7–12°E); Zone 2: nBUS N upwelling cell (19–21°S, 8–14°E); Zone 3: nBUS Center upwelling cell (21–23°S, 9–14°E); Zone 4: nBUS S upwelling cell (23–26°S, 10–15°E); Zone 5: Lüderitz upwelling cell (26–28°S, 10–15°E) and Zone 6: the Orange River–Namaqua upwelling cell (28–30°S, 11–17°E). For the purposes of overlap with the main M. paradoxus distribution, only Zones 4–6 were used.

Average monthly sea surface temperature (SST) values were extracted from the NOAA National Centre for Environmental System (NCEP) database2 (Reynolds et al., 2002) for the area 24–28°S, 13–15°E from January 1982 to October 2016.

Indicators of remote drivers of M. paradoxus growth included the mean monthly multivariate ENSO index (MEI) (Wolter and Timlin, 1993) as an indicator of El Niño–Southern Oscillation (ENSO) as well as the Antarctic Annular Oscillation Index (AAO), which is the dominant 700 mb height anomalies poleward of 20°S that captures decadal-scale climate variability in the Southern Hemisphere (Thompson and Wallace, 2000). These climate indicators were correlated with the BLUP of annual M. paradoxus growth. MEI and AAO were obtained from KNMI Climate Explorer.3

Historical Length-at-Age
Finally, in order to view possible changes in fish growth over time over a period longer than the chronology time series, ALKs were retrieved from literature (ICSEAF, 1977, 1978, 1979, 1980, 1981, 1982, 1983b, 1984, 1985, 1987) for M. paradoxus caught in ICSEAF Divisions 1.3, 1.4 and 1.5 (off Namibia). ALKs from the three divisions and different countries and fleets were pooled by 3-month periods for each year (January–March, April–June, July–September, and October–December). Mean length at each age group a (1–8) was then calculated from each pooled ALK by:

$$\frac{\sum f_i x_i}{\sum f_a}$$

where \( f_i \) = frequency of the age group in length class \( i \), \( x_i \) = class mark of length class \( i \) and \( \sum f_a \) = total frequency in age group \( a \). (Sum of lengths of each individual fish in that age group divided by the total number of fish in the age group). Historical (ICSEAF data) mean lengths-at-age were used together with the current (Supplementary Table S1), and time-related directional changes in fish growth were investigated by linear regressions of mean lengths against year of capture (for each age group).

RESULTS
The best supported model of the growth increment data (Supplementary Figure S2) was that which included fixed effects for Age and Axis, and random effects intercept and slope for Year and FishID (AICc = 603.14, Conditional \( R^2 \) = 0.820, Supplementary Tables S3, S4 and Table 3). The selected model

### Table 3 | Optimal model parameter estimates and test statistics describing Fixed and Random sources of growth variation in Merluccius paradoxus caught along the Namibian coast 1991 to 2013 (Table 1).

| Covariate | Estimate | SE  | t-value |
|-----------|----------|-----|---------|
| Intercept | −2.92    | 0.023 | −127.01 |
| Age       | −0.48    | 0.022 | −21.96  |
| Axis|Long      | 1.24  | 0.041  | 30.33   |

| Covariate | Variance component | SD  | Correlation |
|-----------|--------------------|-----|-------------|
| FishID    | 0.0319             | 0.1786 |             |
| Age| FishID             | 0.0170 | 0.1304 | 0.28 |
| Year      | 0.0023             | 0.0479 |             |
| Residual  | 0.0925             | 0.3042 |             |
| Conditional \( R^2 \) | 0.82 |

Growth modeling was limited to years with ≥3 increment measurements 1982–2013. Random effects age slopes for each individual are denoted by Age| FishID. Axis is the measurement axis on the otolith and is compared with respect to Short. SE is the standard error of the fixed effects covariate estimated as \( s^2/(N−3) \), where \( s^2 \) = the sum square error of each component, and \( N = 834 \). For the random-effects section, the variance component is calculated as \( SD^2 \) (SD, standard deviation, which is estimated), with a correlation of 0.28 between by-FishID slope and intercept.

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1http://disc.sci.gsfc.nasa.gov/Giovanni
2https://iridl.ideo.columbia.edu/SOURCES/NOAA/NCEP/EMC/CMB/ _GLOBAL/.Reyn_SmithOIv2/
3climexp.knmi.nl/
explained 82% of the variation in *M. paradoxus* growth. Age-at-capture did not add significantly to the explanation of otolith growth variation, meaning there was no evidence for biases associated with the sampling scheme. Sex of the fish, cohort, and latitude at which the individual fish was caught did not add significantly to changes in growth variation. Growth declined significantly with age (Figure 3A). The differences in mean increment width against age between short and long axes are shown in Figures 3B,C. From the random effects of Year, annual growth was lowest for 2003, after which mean growth rates appeared to increase toward 2013. However, no significant negative or positive trend in predicted growth was apparent over the time series from 1982 to 2013, which was dominated by decadal variability rather than a long-term trend (Figure 4 and Table 4).

Mean annual population-level growth at age 3 years varied between 6 and 9 cm per year, 2000–2013 (Figure 5). Age-3 mean population-level growth per year was strongly positively correlated with the mean (individual-level) BLUP of growth from otolith increments (\( \rho = 0.730, n = 14, p < 0.01 \)), overlapping from 2000 through 2013.

The GLMM-estimated relative weight (Supplementary Table S5) was used as an estimate of summer fish condition factor of *M. paradoxus* (see Supplementary Figure S3).
Condition factor was weakly positively correlated with BLUP of growth from otolith increments ($\rho = 0.358$, $n = 20$, $p < 0.2$, Figure 6), overlapping in the period 1994–2013.

The best environmental indicators for *M. paradoxus* BLUP of growth from otolith increments were the August and July–September mean upwelling index in the 30°S area, southern Benguela (Supplementary Figure S4B), the August, October, and August–October (austral winter to spring) mean chl-a concentrations at the Namaqua upwelling cell in the southern Benguela (Supplementary Figure S5B) and the October and previous December–February SSTs in the 24–28°S area (Supplementary Figure S6). BLUP of otolith growth also significantly positively correlated with the August upwelling index ($\rho = 0.411$, $n = 32$, $p < 0.05$) and the July–September upwelling index ($\rho = 0.414$, $n = 32$, $p < 0.05$) 1982–2013 at 30°S in the southern Benguela (Figure 7). BLUP of otolith growth also significantly positively correlated with the August ($\rho = 0.734$, $n = 12$, $p < 0.01$), October ($\rho = 0.587$, $n = 12$, $p < 0.05$) and August–October ($\rho = 0.600$, $n = 12$, $p < 0.05$) mean chlorophyll-a concentrations 2002–2013 in the 28–30°S area (Figure 8). Predicted growth significantly negatively correlated with October (austral spring) ($\rho = −0.381$, $n = 32$, $p < 0.05$) and previous December to February (austral summer) ($\rho = −0.355$, $n = 32$, $p < 0.05$) mean SSTs 24–28°S (Figure 9). Mean annual growth did not correlate with the upwelling index at the Lüderitz upwelling cell (Supplementary Figure S4A), the chlorophyll-a concentrations in all other areas (Supplementary Figure S5), the climate data lagged by 1 year (Supplementary Figures S7–S9) or the MEI or the AAO (Supplementary Figure S10).

Trends in mean population-level lengths-at-ages 1–8 from 1977 to 2016 showed that lengths-at-ages 1–4 increased by between 0.02 and 0.08 cm · year$^{-1}$ and mean length-at-ages 5–8 decreased by between 0.004 and 0.25 cm · year$^{-1}$ (Figure 10 and Table 4). Significant time-related changes in mean lengths were seen for age 2 by 0.054 cm · year$^{-1}$ ($t = 2.48$, df $= 42$, $p = 0.0017$), age 3 by 0.075 cm · year$^{-1}$ ($t = 3.04$, df $= 43$, $p = 0.0040$), age 7 by $−0.14$ cm · year$^{-1}$ ($t = −2.13$, df $= 34$, $p = 0.040$), and age 8 by $−0.25$ cm · year$^{-1}$ ($t = −3.58$, df $= 30$, $p = 0.0012$) (Table 4).

**DISCUSSION**

In this study we developed a 32-year otolith chronology for *M. paradoxus*. This is the first study of its kind in the Benguela upwelling system. We showed that otolith chronologies (individual-level growth) as well as ALKs (population-level mean length at age) can be used as indicators of annual fish growth, fish condition and local climatic forcing, despite otoliths of the hakes and other species in the Benguela being notoriously difficult to read. By correlating annual otolith growth with annual fish growth age 3–4 using two independent methods of growth estimation, we indirectly validate age determination of *M. paradoxus* using the whole otolith method and its current age determination criteria. This confirms that *M. paradoxus* grow 6 to 9 cm year$^{-1}$ at 3-years-old (Figure 4), with higher interannual variations (3–12 cm year$^{-1}$) apparent for other ages and in other years (Figure 9). Similarly, Wilhelm et al. (2015a, 2019) calculated growth of northern Benguela *M. paradoxus* using the von Bertalanffy growth curve fitted to 1999 to 2007 pooled lengths at ages 7.3 cm year$^{-1}$ at age group 1 to 4.4 cm year$^{-1}$ at age group 8. The mean growth of 8.32 cm year$^{-1}$ (all ages) that was estimated by a geostatistical model using survey length-frequency distributions of *M. paradoxus* caught off South Africa and Namibia (Jansen et al., 2017) is faster than this, probably driven by South African-caught *M. paradoxus* that grow 10–12 cm year$^{-1}$ until about 50–60 cm length and age 5 years (Durholtz et al., 2015). This means that, compared to *M. capensis* in the northern Benguela.
DF = 30

Otolith increments

### TABLE 4 | Regression results of (A) Mean length-at-age 1–8 years against year of sampling (Year), calculated from otolith-based age-length keys of *Maurolicus paradoxus* caught off Namibia 1977 to 1987 from literature (ICSEAF, 1977, 1978, 1979, 1980, 1981, 1982, 1983b, 1984, 1985, 1987) and 1999 to 2016 from current data (SP, MFM, and Supplementary Table S1) and (B) BLUP of growth from the otolith increments against year.

| Coefficient          | Estimate | SE  | t-value | p-value |
|----------------------|----------|-----|---------|---------|
| **(A) Mean length-at-age x** |          |     |         |         |
| 1-year, DF = 35      |          |     |         |         |
| Intercept            | −14.40   | 60.93 | −0.24   | 0.81    |
| Year                 | 0.019    | 0.031 | 0.61    | 0.55    |
| 2-years, DF = 42     |          |     |         |         |
| Intercept            | −80.37   | 43.55 | −1.85   | 0.072   |
| Year                 | 0.054    | 0.022 | 2.48    | 0.017   |
| 3-years, DF = 43     |          |     |         |         |
| Intercept            | −114.43  | 48.89 | −2.34   | 0.024   |
| Year                 | 0.075    | 0.025 | 3.04    | 0.0040  |
| 4-years, DF = 41     |          |     |         |         |
| Intercept            | −22.52   | 60.26 | −0.37   | 0.71    |
| Year                 | 0.032    | 0.030 | 1.06    | 0.29    |
| 5-years, DF = 39     |          |     |         |         |
| Intercept            | 133.63   | 79.01 | 1.69    | 0.10    |
| Year                 | 0.042    | 0.040 | 1.06    | 0.30    |
| 6-years, DF = 36     |          |     |         |         |
| Intercept            | 184.048  | 97.11 | 1.90    | 0.066   |
| Year                 | −0.064   | 0.049 | −1.32   | 0.19    |
| 7-years, DF = 34     |          |     |         |         |
| Intercept            | 349.56   | 135.20 | 2.59   | 0.014   |
| Year                 | −0.14    | 0.068 | −2.13   | 0.040   |
| 8-years, DF = 30     |          |     |         |         |
| Intercept            | 555.06   | 136.29 | 4.07  | 0.00031 |
| Year                 | −0.25    | 0.068 | −3.59   | 0.0012  |

**(B) Otolith increments**

| Coefficient | Estimate | SE  | t-value | p-value |
|-------------|----------|-----|---------|---------|
| DF = 30     |          |     |         |         |
| Intercept   | 0.0056   | 1.06 | 0.053   | 0.996   |
| Year        | −0.0000027 | 0.00052 | −0.0053 | 0.996 |

Year slopes that are significant at the 95% level are marked with *, and those significant at the 99% level are marked with **. SE is the standard error of each coefficient estimated as $s^2/DF$, where $s^2$, the sum square error of the estimate, and DF, degrees of freedom (N − 2), which is also indicated for each model.

(Wilhelm et al., 2015b, 2017, 2018, 2019), *M. paradoxus* exhibits slower growth, with stable and more consistent annual zone formation on their otoliths.

Otolith growth rates were dominated by high individual and inter-annual growth rate variation, rather than sex-specific or cohort-specific variation in the present study. The inter-annual variability of *M. paradoxus* growth appears to be influenced by local environmental conditions with long-term trends driven by other, possibly fisheries-induced, adaptive changes.

The best predictors for *M. paradoxus* growth were the July–September (austral winter–spring) upwelling index at 30°S and the August–October chl-a concentration in the 28–30°S area, the Namaqua upwelling cell, as well as spring and summer SSTs further northwards of this upwelling cell. Upwelling promotes primary productivity, reflected in the chlorophyll-a concentration, and gives rise to secondary productivity and increased abundances of forage fishes, and in that way promotes growth of predators. *M. paradoxus* at 35–50 cm (ages 3–5 years) mainly feed on krill (euphausiids and macrozooplankton) and myctophids (such as *Lampanyctodes hectoris*) (Wilhelm et al., 2015a) or other mesopelagic fishes such as *Maurolicus muelleri* (Armstrong and Prosch, 1991). Abundances and reproductive activity of these prey populations have been linked with the strength of upwelling (e.g., Hulley and Prosch, 1987; Hutchings et al., 1995; Landaeta and Castro, 2002). This provides evidence that *M. paradoxus* growth is limited by upwelling, i.e., productivity and consequently food availability, reflected by high chl-a concentrations, and cold sea surface temperatures in winter-spring at the Namaqua upwelling cell and slightly northwards. *M. paradoxus* growth also responds positively to cold summers (cold SST in November to February). Summer SST averaged over a large area is mostly a function of the seasonal surface irradiation signal rather than upwelling (Demarcq et al., 2007). Bottom temperatures usually lag SSTs by about 6 weeks (Wilhelm, 2012) and *M. paradoxus* have been associated with bottom temperatures between 6 and 8°C and were most negatively associated with warm (>10°C) bottom temperatures (Mbatha et al., 2019). We thus argue that warm water temperatures in summer would directly negatively affect growth rates of *M. paradoxus* (as opposed to reflecting upwelling productivity) when the increased energy demand generated by warm temperature is not matched with the necessary food supply (e.g., van der Sleen et al., 2018).

Demarcq et al. (2003, 2007) demonstrated, based on SeaWiFS satellite data 1997–2003, that phytoplankton blooms peaked in December to March off Lüderitz (26°S) and October to March off St. Helena Bay (33°S). On a broader spatial scale, a distinct chlorophyll-a minimum could be seen around Lüderitz (26–27°S), the main upwelling cell of the entire Benguela. They also demonstrated that chlorophyll-a was less variable in the southern than in the northern Benguela, but upwelling intensity at the Namaqua upwelling cell (29–31°S) peaked in October–December (Demarcq et al., 2003, 2007). In the present study, *M. paradoxus* growth, however, appeared to be driven by the austral winter–spring production of the Namaqua upwelling cell (28–30°S), where the peak in chl-a concentration was usually around that time (July–October) in recent years (Supplementary Figure S11). The reason for this discrepancy in seasonality of phytoplankton blooms in the Namaqua area between Demarcq et al. (2003, 2007) and values from the present study, is that Demarcq et al. (2003, 2007) used average coastal chl-a concentrations (20 km × 20 km), while the values in the present study were averaged from 11 to 17°E, 28 to 30°S, which stretches beyond the shelf edge to >1,000 m bottom depth (see Figure 1). Integrating chl-a concentrations between coastal and oceanic regions would have resulted in a different seasonality (Lamont et al., 2018b). Mesopelagic fishes are distributed from the coastal area to beyond the shelf edge (Hulley and Prosch, 1987). Their distribution, reproduction and vertical migration (and therefore availability to their predators) is dependent on upwelling productivity (Hulley and Prosch, 1987;
Armstrong and Prosch, 1991), similar to mesopelagic fishes in other upwelling areas (e.g., Landaeta and Castro, 2002). This is a plausible explanation why *M. paradoxus* growth was sensitive to winter upwelling at the Namaqua upwelling cell in the present study. We thus establish that *M. paradoxus* growth variability is a suitable indicator for environmental variability and change, particularly upwelling and primary productivity along the trans-boundary area between the northern and southern Benguela, and between Namibia and South Africa (28–30°S).

Local/small-scale variability in the southern Benguela (as used in the present study) is not necessarily driven by large-scale climatic indices such as the ENSO or AAO. No significant correlations were found between SSTs and AAO by Rouault et al. (2010). According to Tim et al. (2015), El Niño events (ENSO) weakly increase upwelling, and La Niña events (AAO) strengthen upwelling in the southern Benguela. However, correlations between ENSO and upwelling indices were evident only in the summer months, where the strongest upwelling peaks were seen in the southern Benguela (Tim et al., 2015). Since *M. paradoxus* growth is most strongly related to winter upwelling/productivity, this explains why *M. paradoxus* growth was not sensitive to the MEI or AAO (remote forcing) in the present study. Similarly,
the summer upwelling mode in the California current (39°N) (June–August, corresponding to winter time in the Benguela), was not driven by the same forcing mechanism as the winter upwelling mode (sea level pressure and ENSO), but was instead linked with more local and fine-scale gradients and low-frequency processes (Black et al., 2011). Black et al. (2011, 2014) used fish growth from chronologies and fledging success across different taxa to show that some taxa were driven by the summer upwelling mode and some taxa by the winter upwelling mode. This could be the case in the Benguela upwelling system and linkages of growth of different species with winter and summer climate and with different local or remote drivers should be further explored.

*Mero scius paradoxus* growth sensitivity to upwelling variability may have changed over time, having a stronger correlation with upwelling indices in the present than the past. A different growth response to environmental variability is likely to be fishery-induced, similar to what was shown for purple wrasse in Tasmania (Morrongiello et al., 2019). Also, the directional change of fish growth seen in the population-scale data in the present study, is likely to be a fishery-induced adaptive change. Mean lengths at ages 2 and 3 have increased, so cumulative growth up to age 3 years has significantly increased since 1977. Since *M. paradoxus* A50 occurs at 4–6 years of age (Durholtz et al., 2015; Wilhelm et al., 2019), age 3 fish are pre-mature fish. Conversely, length-at-age of the 7–8-year-old post-mature fish have significantly decreased since 1977. All this occurred over a period of 40 years. This speaks to the evolutionary response to fishing, when fishing selection favors increased resource acquisition, faster growth, early maturation, and increased reproductive investment because of reduced longevity (Enberg et al., 2012). This is
visible in the population by increased growth rates up to sexual maturity, smaller post-mature fish (reduced size-at-age) and at the same time decreased length-at-maturity (Enberg et al., 2012). Overall growth rates of adults may also increase due to easing of density-dependent processes related to fishing (Lorenzen and Enberg, 2002; Morrongiello et al., 2019). In terms of fishing on *M. paradoxus* in Namibia, the maximum hake catches (for both species combined) in Namibia occurred in 1971 and after an initial stock decline again in 1985–1986 (Wilhelm et al., 2015a). Hake catch data in Namibia are only separated by hake species from 1997 onwards. However, given that the vessel strength and capacity and therefore fishing depth of the hake-directed fleet has increased since the start of the fishery (early 1960s), and while the *M. capensis* resource became heavily depleted toward 1990 (Kirchner et al., 2012), the proportion of *M. paradoxus* in the hake catches would have substantially increased in that period (1970–1990). From 1997 to current, between 50 and 80% of the total hake catch in Namibia consisted of *M. paradoxus* (Johnsen and Kathena, 2012), while in the 1980s the majority of the hake catch in Namibia was *M. capensis* (van der Westhuizen, 2001), leading to the evolutionary response of fast growth to this increased fishing mortality on *M. paradoxus* seen in the present study.
The results of the present study are in line with results of previous studies on other species and/or in other areas. For example, Duncan (2019) developed a chronology of the resident temperate reef fish *Chrysoblephus laticeps* found along the South African South coast. He negatively correlated annual growth of this species with cumulative intensity of extreme cold-water events and positively correlated annual growth with austral autumn (growing season) water temperature. Duncan (2019) also showed that growth rates of young fish (<12 year-olds) were higher in the exploited population than those in the protected population, and for old (>18-year-old) fish this scenario switched, highlighting a fisheries-induced evolutionary effect similar to the effect described in the present study. In European hake (*M. merluccius*), Vieira et al. (2020) established a relationship between predicted otolith growth and both winter SSTs and spring bottom temperatures. They in fact showed that winter SSTs and spring bottom temperatures affected growth differently in different age groups. They also showed that *M. merluccius* predicted otolith growth was related to recruitment, ascribing this to density dependence. Using another species in the same genus as a species occurring in the Benguela, Atlantic horse mackerel *Trachurus trachurus*, Tanner et al. (2019) identified relationships between otolith growth and SSTs, primary productivity and fish biomass (density dependence). van der Sleen et al. (2018) also revealed density-dependent (SSB and CPUE) as well as temperature-dependent (autumn bottom temperature) changes in otolith growth of European plaice (*Pleuronectes platessa*) in the North Sea. Gillanders et al. (2012) determined strong correlations of otolith growth of a temperate reef fish in the southern hemisphere with summer SSTs.

The predicted climate-induced changes in the Benguela are that upwelling-favorable winds will decrease in the northern Benguela as the pole-ward region of the Benguela, but increase toward the southern Benguela and the Namaqua area (Rykaczewski et al., 2015; Wang et al., 2015). This would negatively affect *M. paradoxus* growth in the northern Benguela as *M. paradoxus* growth was negatively correlated with SST in the northern Benguela (decreased upwelling would mean warmer SST), but conversely would affect *M. paradoxus* growth positively in the southern Benguela. Lamont et al. (2019) determined a general increase in chl-a concentration in the northern Benguela 1997 to 2017, with an increasing upwelling trend in summer.

![Figure 9](image-url)
and winter; and a slight increasing trend in chl-a in the open ocean domain in the southern Benguela. This would have a positive effect on *M. paradoxus* growth rates. In addition to the climate-induced growth rate changes, there are other, age-specific and long-term (most likely fisheries-induced) effects on growth highlighted in the present study, which need to be considered in future population models. Individual-level growth rate affects fish population productivity (e.g., Morrongiello et al., 2014). This is important to take into account for an EAF management of fisheries on the heavily exploited resource, *M. paradoxus*, shared between the northern and southern Benguela. Time-varying growth should be incorporated in stock assessment and prediction models in the future (e.g., JABBA-Select, Winker et al., 2020), and ultimately, together with error estimates, feed into a comprehensive expert system approach as proposed by Jarre et al. (2006). Future models of *M. paradoxus* growth need to incorporate other environmental factors that have not been tested here, such as bottom temperatures and oxygen concentrations. Low oxygen events, which have been shown to affect hake distribution and mortality occur regularly in the northern Benguela (Hamukuaya et al., 1998; van der Lingen et al., 2006a). However, bottom temperatures and oxygen concentrations are not available on such a fine time-scale in the Benguela. It is therefore necessary and important to record temperature, oxygen and salinity regularly, including every bottom trawl station on fisheries-independent surveys, in the future.

The fact that pre-mature *M. paradoxus* growth has increased and post-mature fish growth has decreased shown by the ALK data in the present study, may explain why the overall individual growth rates from the chronology data (averaged over all age groups) show no significant long-term trend. Not enough samples were available for the otolith increment width data in the present study for these growth rates to be separated by age group. As Enberg et al. (2012) discuss, length-at-age data need to also usually be complemented with maturation information. However, L50 data for *M. paradoxus* in Namibia are only available from 1994 onwards, and from 1994 to present L50 remained constant. Age-at-50%-maturity (A50) data are available from 1999 onwards, and has remained constant between 5.5 and 8 years since then (MFMR, unpublished data). It is therefore imperative that age data (and consequently A50 data) are updated from otolith archives at MFMR back to 1990, and analyzed for the southern Benguela. For Cape horse mackerel *Trachurus capensis*, another heavily exploited fish resource in the northern Benguela, L50 decreased by about 2 cm from the 1950s to the early 2000s (van der Lingen et al., 2006b). For sardine *Sardinops sagax* in the Benguela, changes in condition factor (CF), gonad mass and L50 have been linked with density dependence, with CF and gonad mass increasing and L50 decreasing in years of low population density (van der Lingen et al., 2006c).

Apart from adding biological information from existing data on the hakes in the Benguela, there is a further need to investigate species-specific growth responses to environmental forces, especially in the Benguela upwelling system, and with longer time series of otoliths available. These types of analyses on individual species have been used, for example, to detect regime shifts, in the Bering Sea (van der Sleen et al., 2016) and the Baltic Sea (Smoliński and Mirny, 2017) and the Canary Current upwelling system (Tanner et al., 2019). In order to detect a regime shift in otolith growth for the Benguela, there is a need for using a longer otolith increment width time series spanning the time before and across the northern Benguela regime shift.
that occurred during the 1980s (Curry and Shannon, 2004; van der Lingen et al., 2006b). Different taxa have different growth responses to extreme environmental forces and the same species respond differently in different regions (e.g., Izzo et al., 2016; van der Sleen et al., 2016), and even different age groups respond differently for the same species in the same region (e.g., van der Sleen et al., 2018; Vieira et al., 2020). In another example, species across different taxa and trophic levels in the same region, the California Current Upwelling System, were influenced by either winter or summer upwelling modes (Black et al., 2011). Tanner et al. (2019) illustrated that in the Canary Current many different factors, biotic and abiotic factors interacted to influence ecosystem productivity. All of these examples highlight a further need for investigating such responses for different taxa in the Benguela upwelling system (Morrongiello et al., 2012). Otolith chronologies, as well as long term time series of length-at-age of the aquatic resources in the Benguela, such as provided in the present study, prove useful for such long-term (ecosystem-level) indicators and for the possibility to forecast climate- and fisheries-induced changes in growth and ultimately population abundances.

DATA AVAILABILITY STATEMENT

The datasets generated for this study are available on request to the corresponding author.

ETHICS STATEMENT

Ethical review and approval was not required for the animal study because we were not working with any living animals, only otoliths and other secondary data that had been collected for other purposes on research surveys.

AUTHOR CONTRIBUTIONS

BB conceptualized the study. SP acquired and organized the other otoliths and other secondary data that had been collected for other purposes on research surveys. TL, CB, and DL acquired and calculated the upwelling index, SST and chlorophyll datasets, respectively. MW developed the methods for and acquired the otolith chronology data, performed all the statistical analyses, and led the writing of the manuscript. All authors contributed to the manuscript revision, read and approved the submitted version.

REFERENCES

Abramoff, M. D., Magelhaes, P. J., and Ram, S. J. (2004). Image processing with Image. J. Biophotonics Int. 11, 36–42.
Acker, J. G., and Leptoukh, G. (2007). Online analysis enhances use of NASA earth science data. EOS Trans. Am. Geophys. Union 88, 14–17.
Armstrong, M. J., and Prosch, R. M. (1991). Abundance and distribution of the mesopelagic fish Maurolucius muselleri in the southern Benguela system. S. Afr. J. Mar. Sci. 10, 13–28.
Bates, D., Maechler, M., Bolker, B., and Walker, S. (2015). Fitting linear mixed-effects models using lme4. J. Stat. Softw. 67, 1–48. doi: 10.18637/jss.v067.i01
Baudron, A. R., Needle, C. L., Rijnsdorp, A. D., and Marshall, C. T. (2014). Warming temperatures and smaller body sizes: synchronous changes in growth of North Sea fishes. Glob. Change Biol. 20, 1023–1031. doi: 10.1111/gcb.12514.
BCC, (2015). ECOFISH Workshop on Hake and Horse Mackerel Marginal Increment Analysis and Age Determination, 30 November – 11 December 2015. Swakopmund: The Dome Conference Centre.

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SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fmars.2020.00315/full#supplementary-material
Duncombe Rae, C. M. (2005). A demonstration of the hydrographic partition of

Hutchings, L., Pitcher, G. C., Probyn, T. A., and Bailey, G. W. (1995). “The chemical

Durholtz, M. D., Singh, L., Fairweather, T. P., Leslie, R. W., van der Lingen,

Frontiers in Marine Science | www.frontiersin.org

363ñ372.

Sampl. Bull. Int. Commn. SE Atl

3, 15ñ32.

Enberg, K., Jørgensen, C., Dunlop, E. S., Varpe, Ø, Boukal, D. S., Baulier, L., et al.

Izzo, C., Doubleday, Z. A., Grammer, G. L., Barnes, T. C., Delean, S., Ferguson,

Jansen, T., Kainge, P., Singh, L., Wilhelm, M. R., Durholtz, D., Stromme, T.,

ICSEAF (1979). Sampling bulletin. 1979.

Sampl. Bull. Int. Commn. SE Atl

Bull. Am. Meteorol. Soc.

AICcmodavg: Model Selection and Multimodel Inference

Magnuson, J. J., Webster, K. E., Assel, R. A., Bowser, C. J., Dillon, P. J., Eaton, J. G.,

Enberg, K., Jørgensen, C., Dunlop, E. S., Varpe, Ø, Boukal, D. S., Baulier, L., et al.

ICSEAF (1977). Sampling bulletin. 1977.

Sampl. Bull. Int. Commn. SE Atl

Bull.

Afr. J. Mar. Sci.

Demarcq, H., Barlow, R. G., and Shillington, F. A. (2003). Climatology and

Hulley, P. A., and Prosch, R. M. (1987). Mesopelagic fish derivatives in the southern

ICSEAF (1987). Sampling bulletin. 1987.

Sampl. Bull. Int. Commn. SE Atl

Fish.

3, 15ñ32.

Enberg, K., Jørgensen, C., Dunlop, E. S., Varpe, Ø, Boukal, D. S., Baulier, L., et al.

Demarcq, H., Barlow, R., and Hutchings, L. (2007). Application of a chlorophyll

index derived from satellite data to investigate the variability of phytoplankton in the Benguela ecosystem. Afr. J. Mar. Sci. 29, 271–282.

Dorn, M. W. (1992). Detecting environmental covariates of Pacific whiting

Merluccius productus growth using a growth-increment regression model. Fish.

Bull. 90, 260–275.

Duncan, M. I. (2019). A Physiological Study on a Commercial Reef Fish to Quantify

the Relationship Between Exploitation and Climate Change Resilience. Ph. D.

thesis, Rhodes University, Grahamstown.

Duncombe Rae, C. M. (2005). A demonstration of the hydrographic partition of the Benguela upwelling ecosystem at 26°40’S. Afr. J. Mar. Sci. 27, 617–628.

Durholtz, M. D., Singh, L., Fairweather, T. P., Leslie, R. W., van der Lingen, C. D., Bross, C. A. R., et al. (2015). “Fisheries, ecology and markets of South African hake,” in Hake: Biology and exploitation (1st edition), ed. H. Arancibia (Chichester: John Wiley & Sons, Ltd), 38–69.

Enberg, K., Jørgensen, C., Dunlop, E. S., Varpe, Ø, Boukal, D. S., Baulier, L., et al. (2012). Fishing—induced evolution of growth: concepts, mechanisms and the empirical evidence. Mar. Ecol. Prog. Ser. 133, 1–25. doi: 10.1139/h11-044.

Fox, J. (2003). Effect displays in R for generalised linear models. J. Statist. Softw. 8, 1–27.

Fox, J., and Weisberg, S. (2019). An R Companion to Applied Regression, 3rd Edn.

Thousand Oaks, CA: Sage.

Gillanders, B. M., Black, B. A., Meekan, M. G., and Morrison, M. A. (2012). Climatic effects on the growth of a temperate reef fish from the Southern Hemisphere: a biochronological approach. Mar. Biol. 159, 1327–1333. doi: 10.1007/s00227-012-1913-x

Hamukuya, H., O’Toole, M. J., and Woodhead, P. M. J. (1998). Observations of severe hypoxia and offshore displacement of Cape hake over the Namibian shelf in 1994. S. Afr. J. Mar. Sci. 19, 57–59.

Hardman-Mountford, N. J., Richardson, A. J., Agenbag, J. J., Hagen, E., Nykjær, L., Shillington, F. A., et al. (2003). Ocean climate of the South East Atlantic observed from satellite and wind models. Prog. Oceanogr. 59, 181–221.

Hulley, P. A., and Prosch, R. M. (1987). Mesopelagic fish derivatives in the southern Benguela upwelling region. S. Afr. J. Mar. Sci. 5, 597–611.

Hutchings, L., Pitcher, G. C., Probyn, T. A., and Bailey, G. W. (1995). “The chemical and biological consequences of coastal upwelling,” in Upwelling in the Ocean: Modern Processes and Ancient Records, eds C. P. Summerhayes, K. C. Emeis, M. V. Angel, R. L. Smith, and B. Zeitzschel (New York: John Wiley & Sons, Inc.), Vol. 1, 597–611.

Kanamitsu, M., Ebisuzaki, W., Woollen, J., Yang, S.-K., Hnilo, J. J., Fiorino, M., et al. (1997). Potential effects of climate changes on aquatic systems: laurentian large marine ecosystem. Mar. Ecol. Progr. Ser. 169, 159, 1327ñ1333. doi: 10.1007/s00227-012-1913-x

Jarre, A., Hutchings, L., Kirkman, S. P., Kreiner, A., Tchipalanga, P. C. M., Kainge, P., et al. (2015). Synthesis: climate effects on biodiversity, abundance and distribution of marine organisms in the Benguela. Fish. Oceanogr. 24, 122–149.

Johnsen, E., and Kathena, J. (2012). A robust method for generating separate catch time-series for each of the hake species caught in the Namibian trawl fishery. Afr. J. Mar. Sci. 34, 43–53.

Kanamitsu, M., Ebisuzaki, W., Woollen, J., Yang, S.-K., Hnilo, J. J., Fiorino, M., et al. (2002). NCEP-DOE AMIP-II Reanalysis (R2). Bull. Am. Meteorol. Soc. 83, 1631–1643.

Kathena, J. N., Nielsen, A., Thygesen, U. H., and Berg, C. W. (2016). Hake species (Merluccius paradoxus and M. paradoxus) assessment in the Benguela current large marine ecosystem. Environm. Dev. 17, 193–201.

Kirchner, C., Kainge, P., and Kathena, J. (2012). Evaluation of the status of the Namibian hake resource (Merluccius spp.) using statistical catch-at-age analysis. Environm. Dev. 12, 1–52.

Lamont, T., García–Reyes, M., Bogard, S. J., van der Lingen, C. D., and Sydeman, W. J. (2018a). Upwelling indices for comparative ecosystem studies: variability in the Benguela upwelling system. J. Mar. Syst. 188, 8–16. doi: 10.1016/j.jmarsys.2017.05.007

Lamont, T., Brewin, R. J. W., and Barlow, R. G. (2018b). Seasonal variation in remotely-sensed phytoplankton size structure around southern Africa. Remote Sens. Environ. 204, 617–631.

Lamont, T., Barlow, R. G., and Brewin, R. J. W. (2019). Long-term trends in phytoplankton chlorophyll a and size structure in the Benguela Upwelling System. J. Geophys. Res. Letters 124, 1170–1195. doi: 10.1029/2018JC014334

Landaeta, M. F., and Castro, L. R. (2002). Spring spawning and early nursery zone of the mesopelagic fish Maurolicus paradoxus at the coastal upwelling zone off Talcahuano, central Chile. Mar. Ecol. Progr. Ser. 226, 179–191.

Lorenzen, C. K., and Enberg, K. (2002). Density-dependent growth as a key mechanism in the regulation of fish populations: evidence from among-population comparisons. Proc. R. Soc. Lond. B. 269, 49–54. doi: 10.1098/rspb.2001.1853

Magnuson, J. J., Webster, K. E., Assel, R. A., Bowser, C. J., Dillon, P. J., Eaton, J. G., et al. (1997). Potential effects of climate changes on aquatic systems: laurentian great lakes and precambrian shield region. Hydrolog. Processes 11, 825–871.

Mazerolle, M. J. (2019). AICcmodavg: Model Selection and Multimodel Inference Based on (Q) AIC(c). R Package Version 2.2-2. Available at: cran.r-project.org/package=AICcmodavg (accessed December 3, 2019).

Mbatha, F. L., Yemanze, D., Ostrowski, M., Moloney, C. L., and Lipficht, M. R. (2019). Oxygen and temperature influence the distribution of deepwater Cape
hake Merluccius paradoxus in the southern Benguela: a GAM analysis of a 10-year-time-series. Afr. J. Mar. Sci. 41, 413–427.
Morrongiello, J. R., Beatty, S. J., Bennett, J. C., Crook, D. A., Ikedife, D. N., Kennard, M. J., et al. (2011). Climate change and its implications for Australia’s freshwater fish. Mar. Freshw. Res. 62, 1082–1098.
Morrongiello, J. R., Trouth, N. S., and Smith, D. C. (2012). Aquatic bioclimatologies and climate change. Nat. Clim. Change 2, 849–857. doi: 10.1038/nclimate1616
Morrongiello, J. R., Walsh, C. T., Gray, C. A., Stocks, J. R., and Crook, D. A. (2014). Environmental change drives long-term recruitment and growth variation in an estuarine fish. Glob. Change Biol. 20, 1844–1860. doi: 10.1111/gcb.12545
Morrongiello, J. R., and Thresher, R. E. (2015). A statistical framework to explore ontogenetic growth variation among individuals and populations: a marine fish example. Ecol. Monogr. 85, 93–115.
Morrongiello, J. R., Sweetman, P. C., and Thresher, R. E. (2019). Fishing constrains phenotypic responses of marine fish to climate variability. J. Anim. Ecol. 88, 1645–1656. doi: 10.1111/1365-2656.12999
Pilling, G. M., Kirkwood, G. P., and Walker, S. G. (2002). An improved method for estimating individual growth variability in fish, and the correlation between von Bertalanffy growth parameters. Can. J. Fish. Aquat. Sci. 59, 424–432.
Potts, W. M., Henriques, R., Santos, C. V., Munikin, K., Ansero, I., Dufos, F., et al. (2014). Ocean warming, a rapid distributional shift, and the hybridization of a coastal fish species. Glob. Change Biol. 20, 2765–2777. doi: 10.1111/gcb.12612
R Core Team. (2018). R: A Language and Environment for Statistical Computing. version 3.5.0. Vienna: R Foundation for Statistical Computing.
Reynolds, R. W., Rayner, N. A., Smith, T. M., Stokes, D. C., and Wang, W. (2002). An improved in situ and satellite SST Analysis for climate. J. Clim.
Sarkar, D. (2008). Lattice: Multivariate Data Visualization with R. Springer.
Smolínski, S., and Mirny, Z. (2017). Otolith biocronology as an indicator of marine fish responses to hydroclimatic conditions and ecosystem regime shifts. Ecol. Indic. 79, 286–294.
Strømme, T., Lipinski, M. R., and Kainge, P. (2016). Life cycle of hake and likely impact of climate change. J. Fish Biol.
Svimandi, S., and Mirny, Z. (2017). Otolith biocronology as an indicator of marine fish responses to hydroclimatic conditions and ecosystem regime shifts. Ecol. Indic. 79, 286–294.
Tanner, S. E., Vieira, A. R., Vasconcelos, R. P., Dore, S., Azevedo, M., Cabral, H. N., et al. (2019). Regional climate, primary productivity and fish biomass drive growth variation and population resilience in a small pelagic fish. Ecol. Indic. 103, 530–541.
Thompson, D., and Wallace, J. (2000). Annual modes in the extra-tropical circulation. 1. Month-to-month variability. J. Clim. 13, 1000–1016. doi: 10.1038/519392-y
Tim, N., Zorita, E., and Hünicker, B. (2015). Decadal variability and trends of the Benguela upwelling system as simulated in a high-resolution ocean simulation. Ocean Sci. 11, 483–502.
van der Lingen, C. D., Fréon, P., Hutchings, L., Roy, C., Bailey, G. W., Bartholome, C., et al. (2006a). “Forecasting shell processes of relevance to living marine resources in the BCLME,” in Large Marine Ecosystems, Vol. 14, eds V. Shannon, G. Hempel, P. Malanotte-Rizzoli, C. Moloney, and J. Woods (Amsterdam: Elsevier), 309–347.
van der Lingen, C. D., Shannon, L. J., Curry, P., Kreiner, A., Moloney, C. L., Roux, J-P., et al. (2006b). “Resource and ecosystem variability, including regime shifts, in the Benguela Current System,” in Large Marine Ecosystems, Vol. 14, eds V. Shannon, G. Hempel, P. Malanotte-Rizzoli, C. Moloney, and J. Woods (Amsterdam: Elsevier), 147–184.
van der Lingen, C. D., Weatherfield, N. H., Fréon, P., and van der Westhuizen, J. J. (2006c). Density dependent changes in reproductive parameters and condition of southern hake, Scomberomorus commerson, in response to environmental variability. Afr. J. Mar. Sci. 28, 625–636. doi: 10.2989/18142320609504012
van der Sleen, P., Dzaugis, M. P., Gentry, C., Hall, W. P., and Hamilton, V. (2016). Long-term Bering Sea environmental variability revealed by a centennial-length biocronology of Pacific ocean perch (Sebastes alutus). Clim. Res. 71, 33–45.