Development and validation of a multivariable risk score for prolonged length of stay in the surgical intensive care unit

Wesch Conrad\(^a\), Denhaerynck Kris\(^b\), Schaefer Ursi Barandum\(^c\), Siegemund Martin\(^b\), Wehrli Michael\(^b\), Pargger Hans\(^a\), Look Susanne\(^c\)

\(^a\) University Hospital Basel, Department for Anaesthesia, Surgical Intensive Care, Prehospital Emergency Medicine and Pain Therapy, Basel, Switzerland
\(^b\) University of Basel, Department of Public Health, Institute of Nursing Science, Basel, Switzerland
\(^c\) IB Hochschule, Study Center Berlin, Germany

Summary

BACKGROUND: Chronically critical illness is highly relevant in intensive care units, but the definitions in literature vary greatly. The timely detection of prolonged intensive care unit length of stay is crucial for the intensive care unit. We examined the performance of seven predefined predictive factors of prolonged (>20 days) intensive care unit length of stay in adults on the seventh day of stay in intensive care to develop a risk score.

AIM: To develop and validate a risk score for predicting prolonged length of stay in the surgical intensive care unit.

METHODS: This single centre cohort study formed part of an interprofessional research project in one surgical intensive care unit. We examined the performance of seven predefined predictive factors of prolonged (>20 days) intensive care unit length of stay in adults on the seventh day of stay in intensive care to develop a model (n = 304) and validate it (n = 101) using the area under the receiver operating characteristic curve.

RESULTS: Our risk score assigned different points to the following conditions: Charlson Comorbidity Index >2, minimum albumin <20 g/l between days 1 and 7, mechanical ventilation >14 hr on day 7 and the need for parenteral nutrition on day 7. For a validation data set (n = 101), the area under the receiver operating characteristic curve was 0.89 (95% confidence interval 0.770-0.87). At a cut-off value of 100 points, the decision of sensitivity was 88%, the specificity 75%, the positive predictive value 53%, the negative predictive value 95%, and the model fit \(R^2\) 0.40.

CONCLUSIONS: Our model allowed the timely detection of prolonged intensive care unit length of stay with four candidate predictive factors. The timely identification of patients with prolonged intensive care unit length of stay is possible and could influence the person-centred prevention of chronically critical illness and adequate resource allocation. (Trial registration no DRKS 00017073)

Keywords: critical care, length of stay, risk assessment, chronically critical illness, predictive model

Introduction

Improvements in the medical and nursing fields have led to an increasing number of intensive care patients who survive the acute phase of critical illness [1]. After a stay in the intensive care unit (ICU), the majority of patients recover. However, there is also a growing number of so-called chronically critically ill patients [2]. The definitions of chronically critical illness (CCI) in the literature vary greatly regarding characteristic symptoms and criteria [3, 4]. In summary, CCI is a syndrome with significant metabolic, neuroendocrine, neuropsychiatric and immunological dysfunctions [5]. In this illness trajectory, physiological responses during the acute phase of critical illness fail to balance out the induced stress. Physical disorders and mandatory intensive care interventions lead to dysregulation, a loss of resources and consequently to a prolonged stay in the ICU as a main criterion for CCI [6-8].

Of mechanically ventilated patients, 5 to 10% develop CCI, consuming 20 to 40% of the time resources of the ICU [6, 9]. Half of all CCI patients require long-term care combined with a great risk of nursing home admission. In addition, 40% of CCI patients are readmitted to an acute care hospital within the first 6 months following discharge [5].

The in-hospital mortality rate among CCI patients is 20 to 40% [6]; the 1-year mortality rate is 30 to 72% [5]; 12% of them can live independently after 1 year [6].

CCI represents a huge challenge for the affected patients, their families and the inter-professional team at the ICU. Healthcare systems need to pay special attention to CCI patients [8]. This could be strengthened by the systematic identification of patients with a high risk of CCI in the earliest possible stage of its development [10]. Because of differences in the definitions and the limited applicability of the existing predictive models, the timely detection of CCI...
is challenging. Systematic and timely identification of patients with a prolonged ICU length of stay (PICULOS) at a meaningful point in time could be a potential advantage and a first step toward optimal, evidence-based treatment and care planning for CCI patients [5, 6].

The definition of PICULOS as a length of stay (LOS) of more than 20 days is commonly employed in studies to predict a high risk of CCI [11–13].

Several authors have described the predictive factors associated with PICULOS, which can be categorised as occurring before admission (e.g., comorbidities [5, 14]), on the day of admission (e.g., illness severity [15]), and during the acute phase of a critical illness (e.g., presence of a severe infection [3, 4, 6, 16], hypoalbuminaemia [13, 17], respiratory insufficiency [2, 18], and after the first week in ICU (e.g., nutrition problems [11, 18]), consciousness disturbances [19]).

According to Bellar et al. [8], the chronic phase of a critical illness begins on the seventh to tenth day of an ICU stay. This could be a meaningful point in time to predict PICULOS of more than 20 days. This study addressed the following question: how do a set of seven pre-selected factors perform (sensitivity, specificity, predictive values) in predicting PICULOS (>20 d.) in adult surgical ICU patients on the seventh day of their stay (day 7).

Based on a preliminary project, this study aimed to develop and validate a risk score to predict PICULOS in patients of one surgical ICU in order to contribute to the timely identification of patients who are at a high risk of CCI.

### Materials and methods

#### Preliminary project

Prior to this study, an inter-professional team of the surgical ICU at University Hospital, Basel, associated knowledge about CCI with potentially predictive factors, based on a systematic literature review, personal experience and a self-developed conceptual model of allostatic load in CCI patients [20]. In this nurse-led quality development project, we hypothesised that the systematic identification of PICULOS could be a first step toward optimal care planning for CCI patients in this local context.

Through decision making by consensus and after checking the local data, the project team selected seven predictive factors for assessing the risk of PICULOS (>20 d.) on the seventh day of stay (day 7). Table 1 describes the identified predictors and their operationalisation, measurement and statement in terms of literature and context. These factors form the basis for the present study.

#### Design

This single centre cohort study was based on the recommendations of prognosis research [25, 26], which suggest the exploration of different possible predictive factors in combination with elementary clinical information. We used the prognosis research strategy (PROGRESS) framework [27] to achieve the highest possible standard of study quality, design and analysis.

#### Setting

We performed the study as part of the nurse-led quality development project at the surgical ICU at University Hospital, Basel, Switzerland. Annually, the surgical ICU team (22 beds) cares for approximately 2600 adult patients and

---

### Table 1: Aspects of CCI and derivation of factors to predict on day 7 a long surgical ICU Stay (>20 d.).

| Evidence | Operationalisation as study variables | Time and kind of measurement |
|----------|----------------------------------------|------------------------------|
| Comorbidities | Comorbidities affect initial stress response [6, 8]. “Number of pre-existing comorbidities” is predictive for CCI [8], and is associated with PICULOS [3]. The Charlson Comorbidity Index [14] predicts 1-year mortality for a patient in relation to the presence of 22 conditions. | 1. Value of the revised version of Charlson Comorbidity Index Included: all diagnoses on admission day Excluded: admission diagnosis, new diagnoses during ICU stay On admission day Values of the index (0 points = low to 24 points = high), interval scale |
| Illness severity | In combination with medical conditions and organisational characteristics of the ICU, the Simplified Acute Physiology Score SAPS II can be used as a predictive variable of PICULOS [21]. SAPS II [22] predicts hospital mortality based on 17 factors: physiological variables, age, type of admission, and underlying disease variables. | 2. Value of the Simplified Acute Physiology Score SAPS II 24 h after admission Value of index (0 points = low to 163 points = high), interval scale |
| Level of albumin | Hypoalbuminaemia: result of CCI [4, 6], can predict ICU LOS [13, 17] but the results are not unambiguous [5]. | 3. Minimum plasma albumin value in g/l Between admission day and day 7 One value in g/l, interval scale |
| Presence of a severe infection | Infections are a cause [2, 4, 16], a developmental factor [15] and a consequence [6, 8] of CCI. Markers (C-reactive protein, procalcitonin) show insufficient results [23]. | 4. Therapeutic need for anti-infective drugs (antibiotic, antiviral and/or fungalicidal drugs intravenously), preventive administration excluded At any time between admission day and day 7 Answer yes/no, nominal scale |
| Respiratory insufficiency | Extended need for respiration assistance is a hallmark of CCI [2, 6, 11, 16]. A vicious circle develops: serious illness → immobility → muscular dystrophy → need for mechanical ventilation → complications → serious illness | 5. Time of mechanical ventilation, invasive or noninvasive (definition in our study: PEEP >5 mm Hg and pressure support >3 mm Hg, application via tight fitting face masks) On day 7, from 0:00–24:00 Value in hriday 7, interval scale |
| Problems with feeding | Malnutrition [18] and inadequate caloric intake [11] are predictive for CCI. | 6. Main way of feeding (oral, tube feeding, total parenteral nutrition) On day 7 Oral = 1, enteral = 2, parenteral = 3, ordinal scale |
| Consciousness disturbances | The Sedation Agitation Scale SAS [24] is a common sedation assessment scale that has been validated in ventilated and non-ventilated patients in different ICU. Scale: 7 = dangerous agitation, 6 = very agitated, 5 = agitated, 4 = calm and cooperative, 3 = sedated, 2 = very sedated, 1 = unable to rouse | 7. Maximum deviation of the SAS value from the standard value 4 On day 7 Value of index (1–7), interval scale |

CCI = chronically critical illness; ICU = intensive care unit; PEEP = positive end-expiratory pressure; PICULOS = prolonged intensive care unit length of stay; SAPS = Simplified Acute Physiology Score; SAS = Sedation Agitation Scale
covers all surgical and medical disciplines, with a median age of 66 years (interquartile range [IQR] 53–76) and a median LOS of 0.96 days (IQR 0.77–1.87) (Data from the minimal dataset (2016) of the Swiss Society of Intensive Care Medicine [SGI]).

**Data Collection**

Two consecutive datasets were collected: one for the development (n = 304) and the other for the validation (n = 101) of the predictive model (fig. 1). We included all adult patients (≥18 years) with LOS of seven or more consecutive days (≥7d.), between 1 January 2014 and 31 March 2016 for the development dataset (retrospective) and between 1 April 2016 and 31 December 2016 for the validation dataset. The data collection was ongoing in everyday practice, after implementation of the risk score in our practice. Patients who were discharged to another ICU or who died between day 7 and day 20 were excluded.

All data could be collected completely and were entered in an encoded IBM SPSS© Version 22 database. All variables were collected at the highest possible level of measurement (e.g., ratio scale). The file was created by manually obtaining the data from two different, routinely used electronic medical record systems (MetaVision©, ISMed©).

**Variables and measurements**

Candidate predictive factors of PICULOS were identified in the preliminary project (table 1).

Based on scientific literature [8, 11–13] and personal experience, the outcome variable PICULOS was previously dichotomised as “negative” if the LOS was 7 to 20 days (group 1), and “positive” if LOS was >20 days (group 2). Day 1 was the day of admission, regardless of the time of day. Each day of stay counted as a whole day, regardless of the amount of time spent in the unit on the admission and discharge days.

Additionally, patient characteristics (sex, age on admission day, medical discipline: heart, thoracic, traumatology/orthopaedic, visceral, neurosurgery, other) were collected from the medical files.

**Ethics approval and consent to participate**

The process of collection, storage and processing of data was approved by the corresponding ethics committee (Ethikkommission Nordwest- und Zentralschweiz, EKNZ 2016-00948).

Our study involved pre-existing data only (“further use research”). We did not obtain consent to use all of the data based on different reasons, e.g. high morbidity and mortal-
ity rates. Our results may help future CCI patients to recover faster.

**Data analysis**

We analysed all of the variables in the development dataset descriptively, summarising them as measures of central tendency (mean, median) and dispersion (standard deviation [SD], IQR, range).

Predictors were entered into a multiple logistic regression analysis in order to model the probability of the patients staying longer than 20 days, and we retained significant factors only by manual backward deletion, while monitoring the estimates, confidence intervals and inferences of the remaining variables in the model. We also checked for nonlinear relationships using spline functions. In the case of nonlinearity, the threshold values were defined, which were also validated by the literature and practical experience. On the basis of the resulting regression model, a risk score for each patient was calculated by summing the first two digits of the obtained odds ratios (ORs) (multiplied by 10 then rounded up). The diagnostic characteristics of this risk score were explored using a receiver-operating characteristic (ROC), which provided us with the optimal cut-off for deciding whether or not a patient was at risk of an excessively long stay, using the most distant point from the curve to the diagonal.

The developed algorithm was validated using the validation dataset. We additionally checked exploratively whether the previously detected nonlinearities were located at the same values as found in the development set. IBM SPSS® 22 and SAS 9.4 were used for the data analysis.

**Results**

The demographic and clinical characteristics, as well as candidate predictors, of the development sample (n = 304) are presented in **table 2**.

**Regression analysis**

The results of the regression analyses in **table 3** show: (A) an initial model; (B) a model with retained significant variables; and (C) a model with the need for parenteral nutrition on day 7 (yes/no) and dichotomised variables split along the discontinuous relationships found between PICULOS and the variables Charlson Comorbidity Index (>2 points), mechanical ventilation (>14 hr on day 7), and minimum albumin (<20 g/l from days 1–7). The final models revealed mechanical ventilation >14 hr (OR 9.79; 95% CI 4.73–20.27) to be the strongest predictor. The determination coefficient of the final simplified model (R² = 0.36) indicated that it explained as much of the variability as the initial model (R² = 0.37)

The risk scores were derived from the odds ratios and calculated as follows: mechanical ventilation >14 hr on day 7 scored 98 points, the need for parenteral nutrition on day 7 scored 36 points, the lowest albumin concentration <20 g/l between day 1 and day 7 scored 28 points, and a Charlson Comorbidity Index >2 on day 1 scored 23 points. Appendix 1 includes a tool that can be used to enter individual data to calculate the risk of a long ICU stay.

The risk score had an area under the ROC curve of 0.82 (95% CI 0.77–0.87) with regard to the prediction of PICULOS (fig. 2). The point of maximum discriminatory power

---

**Table 2**: Comparison of study variables in Group 1 (LOS 7–20 d, n = 233) and Group 2 (LOS ≥ 21 d, n = 71) in the development sample (n = 304)

| Variable                           | LOS 7-20 d (n = 233) | LOS ≥ 21 d (n = 71) | Odds ratio (95% CI) |
|------------------------------------|----------------------|---------------------|--------------------|
| Sex, male, % (n)                   | 64.8 (151)           | 73.2 (52)           |                    |
| Age in years, median (IQR)         | 69.0 (56.0–77.0)     | 71.0 (61.0–78.0)    |                    |
| LOS in days, median (IQR)          | 10.0 (8.0–13.0)      | 28.0 (25.0–42.0)    |                    |
| Surgical discipline, % (n)         |                      |                     |                    |
| Heart                              | 40.8 (95)            | 46.5 (33)           |                    |
| Thoracic                           | 7.3 (17)             | 11.3 (8)            |                    |
| Traumatology/orthopaedics          | 12.0 (28)            | 9.9 (7)             |                    |
| Vascular                           | 9.9 (23)             | 15.5 (11)           |                    |
| Neurosurgery                       | 16.4 (38)            | 2.8 (2)             |                    |
| Vascular                           | 4.7 (11)             | 5.6 (4)             |                    |
| Other (internal, gynaecology, urology, graft surgery, other) | 9.0 (21) | 8.4 (6) | |

**Predictors**

| Variable                           | LOS 7-20 d (n = 233) | LOS ≥ 21 d (n = 71) | Odds ratio (95% CI) |
|------------------------------------|----------------------|---------------------|--------------------|
| Charlson Comorbidity Index value on day 1, median (IQR) | 2 (0–4) | 3 (2–5) | 1.21 (1.09–1.33) |
| SAPS II value, mean ± SD (range)   | 59.61 ± 15.70 (12–93) | 63.34 ± 16.17 (25–104) | 1.02 (1.00–1.03) |
| Minimum albumin level in g/l days 1–7, mean ± SD (range) | 17.96 ± 4.04 (8–29) | 15.69 ± (4.36, 9–27) | 0.85 (0.80–0.91) |
| Therapeutic need for anti-infective drugs days 1–7, % (n) | 68.20 (159) | 94.40 (67) | 7.80 (2.74–22.19) |
| Main route of feeding, % (n)       |                      |                     |                    |
| Oral (reference)                   | 16.31 (38)           | 5.63 (4)            |                    |
| Gastroenteral                      | 79.40 (185)          | 74.65 (53)          | 0.08 (0.02–0.28) |
| Parenteral                         | 4.29 (10)            | 19.72 (14)          | 0.21 (0.09–0.49) |
| Duration of MV in hours on day 7, median (IQR) | 7 (2–24) | 24 (24–24) | 1.13 (1.09–1.17) |
| SAS value on day 7, % (n)          |                      |                     |                    |
| 1 Unable to rouse                  | 7.70 (18)            | 18.30 (13)          | 0.80 (0.69–0.94)  |
| 2 Very sedated                     | 12.90 (30)           | 23.90 (17)          |                    |
| 3 Sedated                          | 6.00 (14)            | 8.50 (6)            |                    |
| 4 Calm and cooperative             | 24.50 (57)           | 5.60 (4)            |                    |
| 5 Agitated                         | 21.50 (50)           | 19.70 (14)          |                    |
| 6 Very agitated                    | 24.50 (57)           | 22.50 (16)          |                    |
| 7 Dangerous agitation              | 3.00 (7)             | 1.40 (1)            |                    |

CI = confidence interval; IQR = interquartile range; LOS = length of ICU stay; MV = mechanical ventilation; SAS = Sedation Agitation Scale; SD = standard deviation.
derived from the ROC was at 100 points, where the sensitivity was 82% (indicating the proportion of true positives among all of those who stayed for longer than 20 days) and the specificity was 73% (indicating the proportion of true short-stayers among all those staying for fewer than 21 days). The positive predictive value of 48% indicated that, where the risk score >100 points, the chance that the patient actually stayed for longer than 20 days was slightly less than half. The negative predictive value of 93% reflects the chance that those with scores <100 points would indeed stay for 20 days or fewer (table 4). Interestingly, a risk score of 100 points coincided with the start of an over-proportional increase in the probability of a stay exceeding 20 days beyond that score (fig. 3). This warranted dichotomising the risk score using 100 points as a cut-off in order to determine those patients with positive scores. If entered into a logistic regression analysis, this binary variable predicted PICULOS still at an acceptable R² of 0.30.

Validation

The characteristics of the validation dataset (n = 101) are presented in table 5. Application of the risk score to the prediction of PICULOS resulted in an area under the ROC curve of 0.89 (95% CI 0.83–0.96). For confirmatory purposes, we explored whether a nonlinear trend was again present in the risk score, and found that the optimal cut-off at 100 points was identical to that for the development set.

Table 3: Significance values of the steps of logistic regression analyses in the model in development sample (n = 304).

| Tests                                           | Model A (initial model) | Model B (model with dichotomised variables) | Model C (final model) |
|-------------------------------------------------|-------------------------|---------------------------------------------|-----------------------|
| Omnibus test of model coefficients              | χ²=85.985, p <0.001***   | χ² = 83.001, p <0.001***                    | χ² = 82.146, p <0.001***|
| Goodness of fit                                 | 0.372 (Nagelkerke’s R²) | 0.360                                       | 0.357                 |
| Correct allocation                               | 85.9%                    | 80.3%                                       | 84.5%                 |

Continuous or ordinal variables

| C-Index                                         | 5.895, p = 0.015*         | 7.082, p = 0.008**                   | C-Index >2 points, 6.962, p = 0.008** OR 2.345, 95% CI 1.245–4.417 |
| Minimum albumin                                 | 6.263, p = 0.012*         | 8.176, p = 0.004**                   | Minimum albumin <20 g/l, 6.523; p = 0.011* OR 2.788, 95% CI 1.269–6.125 |
| Therapeutic need for anti-infective drugs       | 2.560, p = 0.110          | --                                         | --                     |
| Duration of MV                                  | 27.586, p <0.001***       | 36.062, p <0.001***                   | MV >14h: 37.716, p <0.001*** OR 9.789, 95% CI 4.727–20.273 |
| Main route of feeding                           | 6.399, p = 0.041**        | 6.565, p = 0.038*                     | Parenteral nutrition: 5.959, p = 0.015* OR 3.582, 95% CI 1.286–9.979 |
| SAS                                            | 0.139, p = 0.709          | --                                         | --                     |
| SAPS II                                         | 0.000, p = 0.998          | --                                         | --                     |

Dichotomous variables

| C-Index                                         | 6.962, p = 0.008**         |                                          |                         |
| Minimum albumin                                 | 6.523; p = 0.011*          |                                          |                         |
| Therapeutic need for anti-infective drugs       | 2.788, 95% CI 1.269–6.125  |                                          |                         |
| Duration of MV                                  | 9.789, 95% CI 4.727–20.273 |                                          |                         |
| Main route of feeding                           | 3.582, 95% CI 1.286–9.979  |                                          |                         |

C-Index = Charlson Comorbidity Index; C-Index >2 = Charlson Comorbidity Index >2 points on admission; % CI = confidence interval; minimum albumin = minimum albumin value day 1–7; minimum albumin <20 g/l = minimum albumin value <20 g/l day 1–7; OR = odds ratio; duration of MV = duration of mechanical ventilation day 7; main route of feeding = main route of feeding day 7; parenteral nutrition = parenteral nutrition as the main way of feeding day 7; SAS = Sedation Agitation Scale Wald-Statistics: * p <0.05, ** p <0.01, *** p <0.001

Table 4: Diagnostic characteristics of the clinically validated risk score cut off in development dataset (n = 304)

| Real subdivision to groups | Total |
|----------------------------|-------|
| Group 2 (LOS >20 d)        | 58    |
| Group 1 (LOS 7–20 d)       | 64    |
| Group 2 (LOS >20 d)        | 58    |
| Group 1 (LOS 7–20 d)       | 64    |
| 13                         | 169   |
| 304                        | 304   |

LOS = length of ICU stay Sensitivity = 58/71 = 82%; Specificity = 169/233 = 73%; Positive predictive value = 58/122 = 48%; Negative predictive value = 169/182= 93%. Positive likelihood ratio = + 3.04. Negative likelihood ratio = 0.25

The diagnostic parameters using this cut-off were slightly higher compared to the developmental data (sensitivity 88%, specificity 75%, positive predictive value 53%, negative predictive value 95%). If entered into a logistic regression analysis, this binary variable predicted PICULOS at an R² of 0.40.
Discussion

Based on a regression model, we developed a dichotomised risk score for predicting on day 7 a PICULOS of >20 days (fig. 4). The model included the factors mechanical ventilation for >14 hr and the need for parenteral nutrition on day 7, lowest albumin <20 g/l in the first 7 days and a Charlson Comorbidity Index >2. This was in line with the results of other studies, reporting specific pre-existing diseases [12], hypoalbuminemia [13], a dependence on MV [11], and parenteral nutrition [12] as predictive factors. The score with a sufficient discriminatory ability facilitated the timely identification of patients with PICULOS on day 7. Almost all of the patients with a negative test result were discharged between days 7 and 20.

In our study, mechanical ventilation for more than 14 hr on day 7 was the most influential factor in predicting PICULOS. A positive test result (>100 points) was impossible without mechanical ventilation >14 hr on day 7. This result suggests that even patients who are ventilated for less than 24 hr on day 7 may be at a high risk of PICULOS. The considerable contribution of mechanical ventilation is also confirmed by other studies [11, 13, 15, 18, 28, 29], although our study showed that a combination of several pre-

| Table 5: Comparison of development sample (n = 304) with validation sample (n = 101). |
|---------------------------------|-----------------|-----------------|
|                                | Development (n = 304) | Validation (n = 101) |
| Sex male, n (%)                | 203 (66.8)         | 67 (66.3)         |
| Age in years, median (IQR)     | 70.0 (58.3–77.0)   | 68.0 (53.0–75.5)  |
| LOS in days, median (IQR)      | 12.00 (8.00–19.75) | 14.00 (9.00–18.50) |
| Surgical discipline % (n)      |                    |                  |
| Heart                          | 42.1 (128)         | 31.7 (32)         |
| Thoracic                       | 8.2 (25)           | 7.9 (8)           |
| Traumatology/orthopaedics      | 11.5 (35)          | 11.9 (12)         |
| Visceral                       | 11.2 (34)          | 12.9 (13)         |
| Neurosurgery                   | 13.2 (40)          | 13.9 (14)         |
| Other (internal, gynaecology, urology, graft surgery, other) | 13.8 (42) | 21.9 (22) |
| Affiliation to group 2, % (n)  | 23.36 (71)         | 23.76 (24)        |
| C-Index, median (IQR)          | 2.00 (1.00–4.00)   | 2.00 (1.00–4.00)  |
| C-Index >2 points, % (n)       | 47.37 (144)        | 38.61 (39)        |
| Minimum albumin level in g/l, mean ± SD (range) | 17.43 ± 4.08 (8–29) | 18.97 ± 4.02 (11–31) |
| Minimum albumin <20 g/l, % (n) | 68.09 (207)        | 63.37 (64)        |
| Duration of MV, median (IQR)   | 12.00(3.00–24.00)  | 8.00(3.00–24.00)  |
| MV >14 hr on day 7, % (n)      | 46.71 (142)        | 47.52 (48)        |
| Parenteral nutrition, % (n)    | 7.89 (24)          | 4.95 (5)          |

C-Index = Charlson Comorbidity Index; IQR = interquartile range; LOS = length of ICU stay; MV = mechanical ventilation; SD = standard deviation
dictive factors has a higher predictive value than the factor mechanical ventilation >14 hr alone.

Significantly more patients with LOS >20 days needed parenteral nutrition. The need for parenteral nutrition was also a predictive factor in the model of Chen et al. [12]. The high demand for enteral nutrition was striking in all patients with an ICU stay >7 days. Parenteral nutrition seems to be a sign of intestinal absorption failure, prohibiting protein anabolism.

A minimum albumin value <20 g/l between days 1 and 7 was found to be a significant predictive factor. The chosen operationalisation proved to be highly suitable. The mere presence of hypoalbuminaemia (albumin value <34 g/l [30]) as a predictive factor would have been insufficient, since all patients with LOS ≥7 days in our study had an albumin value below this defined threshold. Lee et al. [31] also showed that, amongst other things, plasma albumin predicted ICU LOS in general surgery, but it must be kept in mind that albumin levels may change as a result of the infusion of albumin. This lies beyond the scope of our present study.

Our results confirmed the importance of specific comorbidities in determining outcomes following critical illness [32, 33], and the association between chronic comorbidities and PICULOS [3, 5]. The updated version of the Charlson Comorbidity Index [14], originally prepared to predict mortality within 1 year after hospital discharge, is also an appropriate risk factor to identify PICULOS.

The factors SAPS II, therapeutic need for anti-infective drugs, and SAS did not contribute significantly to the prediction of PICULOS. In the case of SAPS II, this may indicate that a very long ICU stay does not depend significantly on the specific acute critical illness. The therapeutic need for anti-infective drugs between days 1 and 7 was also not affirmed as an independent predictor. This could be because a large proportion of the sampled patients (74.3%) needed anti-infective drugs. In the literature, infections are often used as a variable for predicting LOS in ICU [12, 15], notwithstanding the fact that it is unclear whether it is a cause [2, 16], a developmental factor [15], or a consequence [6, 8] of CCI. Likewise, despite finding low SAS values (1–2) for patients with PICULOS >20 days, no independent relationship could be found in an analysis controlled for the variable time of MV, the reason for which is unclear. Brain dysfunction, as well as cognitive symptoms including delirium and memory gaps, are described as typical criteria for CCI [5, 6, 10, 34]. However, the exact cause-and-effect relationship between disturbances in consciousness and the development of CCI remains to be investigated.

In this study, we used PICULOS as an operationalisation of CCI. This outcome variable enabled us to include patients with different diagnoses and treatments. In addition, we were able to form study groups for the group comparisons (see fig. 1). Our open research method emphasises literature recommendations, because PICULOS includes various criteria of CCI, such prolonged mechanical ventilation (PMV) [2].

The inclusion of patients with LOS ≥7 days and the assessment on day 7 seem to be advantageous. This group differed significantly from many factors in the entire ICU population (see table 2). Widyastuti et al. [35] were not able to predict a long ICU stay in individual patients on the basis of assessment on day 1, because most patients had short ICU stays (75th percentile: 1 day). This result appears to be relevant in our unit, where patients also have a low average LOS. The assessment on day 7 allows the inclusion of factors that reflect the acute phase of illness,
such as hypoalbuminaemia. Thus, it is possible to recognise systematically patients at risk, at the beginning of the chronic phase.

Limitations
PICULOS and the corresponding predictors depend on local processes and treatment strategies; for instance, ICUs that use more albumin infusions to increase the serum levels of albumin near 30 g/l will be unable to use our dichotomised score [36]. Experts from our ICU selected the predictive factors through a consensus process, focussing on data available from patient records. This contained a certain degree of subjectivity and might limit the generalisability of our results. Factors that could not simply be derived from the existing documentation or factors that are very difficult to operationalise may have been falsely excluded: for example, sociodemographic factors, such as resilience or family support. We operationalised our outcome variable PICULOS as a dichotomous variable of LOS of more than 20 days. This artificially determined threshold could disadvantage patients who fail fully to meet this predefined condition but who are nevertheless at a high risk of CCI (e.g., a patient discharged from the ICU to an intermediate care unit after 18 days). Riley et al. [25] recommend analysing continuous factors using their continuous scale. They also recommend a prospective rather than a retrospective design, as this produces clear inclusion criteria, more complete baseline and follow-up data, as well as a greater standardisation of the diagnostic and therapeutic procedures. However, there were no missing data in our study, we collected the data directly from the individual patients’ documentation, and we discussed special cases within the study team. Despite these limitations, the risk score works in our setting appropriately, and our approach to develop a local risk score can easily be adapted to other ICUs.

Conclusions for clinical practice
This study marked an important step toward equipping the involved health professionals with an extended understanding of PICULOS, the trajectory of CCI and prognosis research as intrinsic aspects of clinical care [27]. Our model is suitable for systematic application within our ICU. It is important that all responsible health professionals are informed of any positive scores (e.g., during inter-professional ward rounds). To improve the quality of clinically collected data, electronic documentation systems could assist the automatic calculation of the score. We have already successfully implemented this within our electronic documentation. The transferability of our results into comparable settings still has to be investigated.

More objective estimated probabilities can supplement the clinical reasoning and decision making of health professionals [26]. However, it must be kept in mind that the predictive models form only part of good qualitative reasoning. Patients who are at increased risk of CCI (and not only PICULOS) urgently require a comprehensive assessment and patient-centred treatment planning.

The study did not investigate whether the risk score alone affects the development or trajectory of CCI. However, based on the risk score, we intensified the care for patients at risk of PICULOS in our ICU while integrating all therapeutic professionals, tailored family information, systematic communication and coordination, and the development of an evidence-based assessment and treatment plan. From this, one can infer that the assessment based on our model could be the first step toward the optimal treatment and care of patients with a high risk of CCI.

Implications for further research
Updating and advancing a model by exploring additional prognostic factors is often desirable [37]. According to other studies, the presence of pressure sores [5, 11, 13] or the patients’ physical capacity before critical illness [38] could be suitable additional factors. Investigation of psychosocial factors (e.g., resilience, marital status, social support) might also prove rewarding.

Another implication for research is the examination of the clinical impact of the tool on decision-making and patient outcomes [39]. In a comparative study, one patient group with usual care should be compared with another group in which the model’s predictions are made available to health professionals to guide their treatment decisions [37].

Availability of data and materials
The dataset generated and analysed during the current study is not publicly available because the study was an internal quality development project. With the approval of the ethics committee (Ethikkommission Nordwest- und Zentralschweiz EK NZ), the patient data were used without the consent of the patients, as this was disproportional for this purpose. The dataset is available from the corresponding author on reasonable request.

Acknowledgements
Allison Dwilecki, scientific secretary in the Department of Anaesthesiology, University Hospital Basel: editing of the publication text in English. Helmut Wesch, English teacher: Editing of the publication text in English. Assessment of the text as a nonprofessional.

Financial disclosure
This research received no specific grant from any funding agency in the public, commercial, or not-for-profit sector.

Potential competing interests
The authors declare that they have no competing interests.

References
1. Rosseau S, Suttorp N. Der chronisch kritisch kranke Patient [The chronically ill patient]. Med Klin Intensivmed Notf Med. 2013;108(4):266. In German. doi: http://dx.doi.org/10.1007/s00065-012-0162-6.
2. Carson SS. Definitions and epidemiology of the chronically critically ill. Respir Care. 2012;57(6):848–56, discussion 856–8. doi: http://dx.doi.org/10.4187/respcare.01736. PubMed.
3. Kahn JM, Le T, Angus DC, Cox CE, Hough CL, White DB, et al.; ProVent Study Group Investigators. The epidemiology of chronic critical illness in the United States. Crit Care Med. 2015;43(2):282–7. doi: http://dx.doi.org/10.1097/CCM.0000000000000710. PubMed.
4. Macintyre NR, Epstein SK, Carson S, Scheinorn D, Christopher K, Muldoon S; National Association for Medical Direction of Respiratory Care. Management of patients requiring prolonged mechanical ventilation: report of a NAMDRC consensus conference. Chest. 2005;128(6):3937–54. doi: http://dx.doi.org/10.1378/chest.128.6.3937.
5. Wieneke C, Winkelman C. Chronic critical illness: prevalence, profile, and pathophysiology. AACN Adv Crit Care. 2010;21(1):44–61, quiz 63. doi: http://dx.doi.org/10.1097/01.NCC.0000381328.1816e6a2. PubMed.
6. Nelson JE, Cox CE, Hope AA, Carson SS. Chronic critical illness. Am J Respir Crit Care Med. 2010;182(4):446–54. doi: http://dx.doi.org/10.1164/rccm.201002-0210CL. PubMed.
7 Beckie TM. A systematic review of allostatic load, health, and health disparities. Biol Res Nurs. 2012;14(4):311–46. doi: http://dx.doi.org/10.1177/1099801412455688. PubMed.

8 Bellar A, Kankir K, Burkett M. Understanding, recognizing, and managing chronic critical illness syndrome. J Am Acad Nurse Pract. 2009;21(11):571–8. doi: http://dx.doi.org/10.1111/j.1745-7992.2009.00451.x. PubMed.

9 Rodríguez Villar S, Barrientos Yuste RM. Long-term admission to the intensive care unit: cost-benefit analysis of care Exp Anest Reanim. 2014;61(9):489–96. doi: http://dx.doi.org/10.1016/j.redar.2014.02.008. PubMed.

10 Jeitziner MM, Massarutato P, Barudan Schäfer U. Symptombelumst und entsprechende Interventionen. Intensiv. 2015;22(03):123–7. In German. doi: http://dx.doi.org/10.1055/s-0035-1550608. PubMed.

11 Loos SH, Marchese CB, Boniatti MM, Wawrzeńczak IC, Oliveira RP, Nunes LN, et al. Prediction of chronic critical illness in a general intensive care unit. Rev Assoc Med Bras (29). 2012;39(3):241–7. doi: http://dx.doi.org/10.1590/S0104-42302012005300017. PubMed.

12 Chen HY, Varens DI, Golemanian E. A simplified score for transfer of patients requiring mechanical ventilation to a long-term care hospital. Am J Crit Care. 2011;20(6):e122–30. doi: http://dx.doi.org/10.4037/ajcc20111775. PubMed.

13 Snabozki CR, Tellez A, Klika AK, Xu M, Kattan MW, Guzman JA, et al. Predicting discharge to a long-term acute care hospital after admission to an intensive care unit. Am J Crit Care. 2014;23(4):e46–53. doi: http://dx.doi.org/10.4037/ajcc2014985. PubMed.

14 Quan H, Li B, Couris CM, Fushimi K, Graham P, Hider P, et al. Updating and validating the Charlson comorbidity index and score for risk adjustment in hospital discharge abstracts using data from 6 countries. Am J Epidemiol. 2011;173(6):67–82. doi: http://dx.doi.org/10.1093/aje/kwq433. PubMed.

15 Higgins TL, McGee WT, Steingrub JS, Rapoport J, Lemoshow S, Teres D. Early indicators of a prolonged intensive care unit stay: impact of illness severity, physician staffing, and pre-intensive care unit length of stay. Crit Care Med. 2003;31(1):45–51. doi: http://dx.doi.org/10.1097/00003188-200301000-00007. PubMed.

16 Boniatti MM, Friedman G, Castilho RK, Vieira SR, Fialkow L. Characteristics of chronically critically ill patients: comparing two definitions. Clinics (São Paulo). 2011;66(44):701–4. doi: http://dx.doi.org/10.1590/S1807-59322011000400027. PubMed.

17 Lee JH, Waak K, Grosse-Sundrup M, Xue F, Lee J, Chipman D, et al. Global muscle strength but not grip strength predicts mortality and length of stay in a general population in a surgical intensive care unit. Phys Ther. 2012;92(12):1546–55. doi: http://dx.doi.org/10.2522/ptj.j20110403. PubMed.

18 Estenssoro E, Reina R, Canales HS, Saenz MG, Gonzalez FE, Aprea MM, et al. The distinct clinical profile of chronically critically ill patients: a cohort study. Crit Care. 2006;10(3):R89. doi: http://dx.doi.org/10.1186/cc4941. PubMed.

19 Marchioni A, Fantini R, Anetona F, Clini E, Fabbrini L. Chronic critical illness: the price of survival. Eur J Clin Invest. 2015;45(12):1431–9. doi: http://dx.doi.org/10.1111/eci.12587. PubMed.

20 Weesch C. Master Thesis: Developing and validating a multivariable tool to predict a prolonged length of stay in the ICU: a retrospective exploratory cohort study, in Institute of Nursing Science. 2017, University of Basel.

21 Higgins PA, Daly BJ, Lipson AR, Guo SE. Assessing nutritional status in chronically critically ill adult patients. Am J Crit Care. 2006;15(2):166–76, quiz 177. PubMed.

22 Le Gall JR, Lemoshow S, Saulnier F. A new Simplified Acute Physiologic Score (SAPS II) based on a European/North American multicenter study. JAMA. 1993;270(24):2957–63. doi: http://dx.doi.org/10.1001/jama.1993.03510240006035. PubMed.

23 Cox CE. Persistent systemic inflammation in chronic critical illness. Respir Care. 2012;57(6):859–64, discussion 864–6. doi: http://dx.doi.org/10.4187/respcare.01719. PubMed.

24 Riker RR, Picard JT, Fraser GL. Prospective evaluation of the Sedation-Agitation Scale for adult critically ill patients. Crit Care Med. 1999;27(7):1325–9. doi: http://dx.doi.org/10.1097/00003166-199907000-00002. PubMed.

25 Riley RD, Hayden JA, Steyerberg EW, Moons KG, Abrams K, Kyzas PA, et al.; PROGRESS Group. Prognosis Research Strategy (PROGRESS) 2: prognostic factor research. PLoS Med. 2013;10(2). doi: http://dx.doi.org/10.1371/journal.pmed.1001380. PubMed.

26 Moons KGM, Kengne AP, Woodward M, Royston P, Vergouwe Y, Altman DG, et al. Risk prediction models: I. Development, internal validation, and external validation. JAMA. 2012;98(4):683–90. doi: http://dx.doi.org/10.1136/heartjnl-2011-301246. PubMed.

27 Hemingway H, Croft P, Perel P, Hayden JA, Abrams K, Timmis A, et al.; PROGRESS Group. Prognosis research strategy (PROGRESS) 1: a framework for researching clinical outcomes in chronic disability. J Clin Epidemiol. 2013;66(6):e46–e55. doi: http://dx.doi.org/10.1016/j.icejml.2013.04.001.

28 Wein OY, Birnbbaum J, Wennecke K, England M, Konertz W, Spies C. Prolonged intensive care unit stay in cardiac surgery: risk factors and long-term survival. Ann Thorac Surg. 2006;81(3):880–5. doi: http://dx.doi.org/10.1016/j.athoracsur.2005.10.024. PubMed.

29 Vincent JL, Dupuis MJ, Navickis RJ, Wikles MM. Hypoalbuminemia in acute illness: is there a rationale for intervention? A meta-analysis of cohort studies and controlled trials. Ann Surg. 2003;237(3):319–34. doi: http://dx.doi.org/10.1097/00000543-200303000-00004. PubMed.

30 Lee JH, Kim J, Kim K, Jo YH, Rhee J, Kim YW, et al. Albumin and C-reactive protein: predictors of mortality in critically ill patients with community-acquired pneumonia. J Crit Care. 2011;26(3):278–87. doi: http://dx.doi.org/10.1016/j.jcc.2010.10.007. PubMed.

31 de Groot V, Beckerman H, Lankhorst GJ, Bouten LM. How to measure and access the incremental value of a new J Clin Epidemiol. 2005;56(3):221–9. doi: http://dx.doi.org/10.1016/S0895-4356(02)00585-1. PubMed.

32 Rochon PA, Katz JN, Morrow LA, McGlinchey-Berroth R, Ahlquist MM, Sarkarati M, et al. Comorbid illness is associated with survival and length of stay in critically ill patients: comparing two definitions. Clinics (São Paulo). 2011;66(44):701–4. doi: http://dx.doi.org/10.1590/S1807-59322011000400027. PubMed.

33 Lieu J, Kim J, Kim K, Jo YH, Rhee J, Kim YW, et al. Albumin and C-reactive protein: predictors of mortality in critically ill patients with community-acquired pneumonia. J Crit Care. 2011;26(3):278–87. doi: http://dx.doi.org/10.1016/j.jcc.2010.10.007. PubMed.

34 Agitation Scale for adult critically ill patients. Respirology. 2008;13(6):764–76. doi: http://dx.doi.org/10.1111/j.1440-1843.2008.01178.x. PubMed.

35 Couris CM, Picard RL, Moineddin R, Couris R, Moineddin R, Couris R, et al. Prospective evaluation of the Sedation-Agitation Scale for adult critically ill patients. Crit Care Med. 2006;34(2):449–55. doi: http://dx.doi.org/10.1097/01.CCM.0000183177.04124.87. PubMed.
Appendix 1

Tool to calculate the risk of a long ICU stay

The appendix is available as a separate file for downloading at https://smw.ch/en/article/doi/smw.2019.20122/