Decaying or annihilating dark matter particles could be detected through gamma-ray emission from the species they decay or annihilate into. This is usually done by modelling the flux from specific dark matter-rich objects such as the Milky Way halo, Local Group dwarfs, and nearby groups. However, these objects are expected to have significant emission from baryonic processes as well, and the analyses discard gamma-ray data over most of the sky. Here we construct full-sky templates for gamma-ray flux from the large-scale structure within ~200 Mpc by means of a suite of constrained N-body simulations (CSiBORG) produced using the Bayesian Origin Reconstruction from Galaxies algorithm. Marginalising over uncertainties in this reconstruction, small-scale structure, and parameters describing astrophysical contributions to the observed gamma-ray sky, we compare to observations from the Fermi Large Area Telescope to constrain dark matter annihilation cross sections and decay rates through a Markov Chain Monte Carlo analysis. We rule out the thermal relic cross section for s-wave annihilation for all $m_X \lesssim 7 \text{GeV}/c^2$ at 95% confidence if the annihilation produces gluons or quarks less massive than the bottom quark. We infer a contribution to the gamma-ray sky with the same spatial distribution as dark matter decay at $3.3 \sigma$. Although this could be due to dark matter decay via these channels with a decay rate $\Gamma \approx 1 \times 10^{-26} \text{s}^{-1}$, we find that a power-law spectrum of index $p = -2.75_{-0.46}^{+0.71}$, likely of baryonic origin, is preferred by the data.

I. INTRODUCTION

The actual particle nature of dark matter (DM) is not yet known, despite it having 5 times the average density of baryonic matter. Theoretically favourable candidates are weakly interacting massive particles (WIMPs) [1–3], including, but not limited to, the lightest supersymmetric particle in supersymmetric theories. If these WIMPs are thermal relics of the early Universe, then the current abundance of DM suggests that the self-annihilation cross section should be $\langle \sigma v \rangle_{\text{th}} \approx 3 \times 10^{-26} \text{cm}^3\text{s}^{-1}$ [4] (the thermal relic cross section), which is suspiciously similar to what one would expect for a particle of mass $\sim 0.1 - 1 \text{TeV}$ with a coupling comparable to the electroweak coupling of the Standard Model (SM). Recent experiments find an anomalous muon magnetic moment, as measured by the “Muon $g - 2$” experiment [5], and a mass for the $W$ boson that is higher than expected [6]. This further motivates probing physics beyond the SM and thus the search for DM candidates. Detection of such particles could therefore solve some of the fundamental questions of particle physics and cosmology.

DM annihilation or decay for particles at these masses could be detected through the emission of gamma rays by their decay products. Despite the theoretically low interaction rates - one would expect only a few in $10^{45}$ particles to annihilate per Hubble time in the present Universe [7] - the vast quantities of DM on cosmological scales makes these processes potentially detectable in the state-of-the-art gamma-ray measurements from the Fermi Large Area Telescope (Fermi-LAT) [8]. Indeed, the excess of observed gamma rays toward the galactic centre (Galactic Centre Excess; GCE) [9–11] can be fitted well by the annihilation of DM [12–18]. There is debate over whether other explanations could also explain the emission. Several studies [19–24] argue that the GCE can be explained by a population of unresolved point sources, although this is contested [17, 25–27], and some groups [28–31] find both DM annihilation and other models can fit the data. A combination of these two processes is, of course, possible, and is plausible given the spatial variation of the GCE [32]. Other astrophysical processes at the centre of the galaxy [33–38] could also be responsible.

Given these conflicting explanations of the GCE, in order to unambiguously detect or rule out DM annihilation or decay models, one should determine if an excess of gamma rays is detected from other sources or across the full sky. Previous studies have placed constraints on $\langle \sigma v \rangle$ through cross-correlation between Fermi-LAT data
and galaxy [e.g. 39, 40] or lensing [41] catalogues, or by studying nearby dwarf galaxies [42–44] or groups [45]. Moreover, such emission should increase the kinetic energy of baryons, so by considering the impact on the Cosmic Microwave Background (CMB) [46] or galaxy formation [47], one can rule out a velocity-independent cross section for thermal relic DM particles less massive than \(\sim 30\) GeV.

Instead of focusing on a few nearby or massive objects, the aim of this work is to search for the signature of DM decay and annihilation across the full sky by forward-modelling the observed gamma-ray sky, as first suggested in [48]. As proposed in [49], anisotropies in the cosmic gamma-ray background could be a signature of DM annihilation or decay. This has previously been studied through the two-point correlation function [e.g. 50], which is calibrated with unconstrained \(N\)-body simulations [e.g. [51] use the Millennium-II and Aquarius simulations [52–54]). Instead, we utilise the CSiBORG suite of constrained \(N\)-body simulations [55–58]. The initial conditions (ICs) for these simulations are chosen to produce final three-dimensional DM density fields that are consistent with the observed positions of galaxies in the 2M++ galaxy catalogue. The ICs are inferred using the BORG (Bayesian Origin Reconstruction from Galaxies) algorithm [59–63], a fully Bayesian forward model. We use the particle positions in the simulations to produce maps of the expected gamma-ray flux from halos down to \(4.38 \times 10^{14}\) M\(_\odot\) in mass (although we also model smaller substructures), as well as from DM not identified as belonging to halos. The halos are assumed to have Navarro-Frenk-White (NFW) profiles [64] and we explicitly model unresolved substructure in a probabilistic manner, since the signal from DM annihilation is sensitive to the peaks in the density field.

We do not include a contribution from the Milky Way halo or Local Group dwarf galaxies (which are unresolved in our simulations) in our templates so as to produce constraints entirely from large scale structure. These constraints will be free from many of the systematics affecting searches in particular objects and will reveal the amount of constraining power for DM annihilation and decay to be found in various parts of the cosmic web. To account for non-DM effects we include templates for emission due to point sources, galactic emission and an isotropic background. We marginalise over the amplitudes of these templates, as well as the realisations in the CSiBORG suite which sample the full BORG posterior in ICs of the 2M++ volume. We then compare to Fermi-LAT observations via a Markov Chain Monte Carlo (MCMC) algorithm. A full-sky field-level inference allows us to capture not only the two-point statistics but implicitly all higher orders, too.

In this work we rule out the thermal relic cross section at 95\% confidence for annihilations that produce gluons or quarks less massive than the bottom quark if DM has a mass \(m_\chi \lesssim 7\) GeV/c\(^2\). We find a contribution to the gamma-ray sky with the same spatial distribution as expected from DM decay (flux proportional to DM density) at 3.3\sigma confidence, with a decay rate \(\Gamma \approx 6 \times 10^{-28}\) s\(^{-1}\) for these channels. However, a power-law spectrum with an index \(p = -2.75^{+0.71}_{-0.74}\) provides a better fit to the data, suggesting a non-DM origin. In our fiducial analysis, we do not rule out the thermal relic annihilation cross section at any mass for production of top or bottom quarks; we obtain upper bounds that are half the size (i.e. tighter constraints) if we marginalise over the contribution proportional to the DM density, but we do not include this contribution in the fiducial analysis. Our constraints on DM decay to leptons are approximately an order of magnitude less stringent than decay to quarks.

This paper is structured as follows. We discuss DM annihilation and decay models in Section II and introduce the gamma-ray data that we use to constrain these models in Section III. Our inference and template construction methods are outlined in Section IV. The results are presented in Section V and discussed in Section VI, including the potential systematic uncertainties and a comparison to the literature. We conclude in Section VII. Equations throughout the paper use units \(\hbar = c = 1\).

## II. THEORETICAL BACKGROUND

We start by assuming that DM is made of a single particle, \(\chi\), of mass \(m_\chi\), whose antiparticle is itself (e.g., Majorana fermions). This particle is assumed to be able to both decay

\[
\chi \rightarrow A\bar{A},
\]

and annihilate

\[
\chi\chi \rightarrow A\bar{A},
\]

to a standard model particle, \(A\), and its antiparticle, \(\bar{A}\). The annihilation of the produced particles would lead to gamma-ray emission at some energy \(E_\gamma\), which one could detect. If these processes occur via channel \(i\), the photon flux for annihilation per unit density squared at redshift \(z\) is [65]

\[
\frac{d\Phi_{\gamma \gamma}^{\text{ann}}}{dE_\gamma} = \langle \sigma v \rangle \frac{8\pi m_\chi^2}{3} \sum_i \frac{dN_{\gamma \gamma}^{\text{ann}}}{dE_\gamma} \bigg|_{E_\gamma = E_{\gamma,i}(1+z)},
\]

and for decay per unit density

\[
\frac{d\Phi_{\gamma \gamma}^{\text{dec}}}{dE_\gamma} = \frac{\Gamma}{4\pi m_\chi^2} \sum_i \frac{dN_{\gamma \gamma}^{\text{dec}}}{dE_\gamma} \bigg|_{E_\gamma = E_{\gamma,i}(1+z)},
\]

where \(\langle \sigma v \rangle\) is the thermally averaged cross section, \(\tau = 1/\Gamma\) is the lifetime of the particle, and \(dN_{\gamma}/dE_\gamma\) is the photon energy distribution for channel \(i\). Throughout this work we assume s-wave annihilation so the parameter \(\langle \sigma v \rangle\) is assumed to be a constant, i.e. independent of \(v\). If \(\chi\) is not its own antiparticle (e.g., Dirac fermions),
the annihilation flux is half of this value, provided there is no matter-antimatter asymmetry. Since we do not a priori know which branching ratios to use, in this work we assume that the annihilation or decay occurs via a single channel; however, our analysis can be trivially extended to multiple channels.

Since these results apply at unit density, we must now take into account the integrated density along the line of sight. By introducing the $J$ factor

$$\frac{dJ}{d\Omega} = \int \rho_{\text{DM}}^2(s,\Omega) ds,$$

where $s$ is a radial coordinate, and the $D$ factor

$$\frac{dD}{d\Omega} = \int \rho_{\text{DM}}(s,\Omega) ds,$$

we arrive at the total photon fluxes per unit solid angle

$$\frac{d^2\Phi_{\text{ann}}}{dE\,d\Omega} = \frac{d\Phi_{\text{ann}}^{PP}}{dE\,d\Omega} dJ,$$

$$\frac{d^2\Phi_{\text{dec}}}{dE\,d\Omega} = \frac{d\Phi_{\text{dec}}^{PP}}{dE\,d\Omega} dD,$$

where we note that we have assumed that the cosmological redshift variation across the source is negligible, so we can factor out Eqs. 3 and 4 from the line of sight integral.

III. GAMMA-RAY DATA

In this work we use gamma-ray observations from Fermi-LAT and analyse these using the Fermi Tools\(^1\) and FermiPy \([66]\). To mitigate the effect of contamination from cosmic rays we consider photons within the event class SOURCEVETO. We select all photons in this event class of energy $500\,\text{MeV} - 50\,\text{GeV}$ between mission weeks 9 and 634 which are flagged as belonging to the upper quartile of angular resolution (PSF3) and set the maximum zenith angle to be 90°. We subdivide these data into nine logarithmically space energy bins, and then bin spatially onto HEALPix\(^2\) \([67, 68]\) maps. Although the angular resolution of the data corresponds to $n_{\text{side}} \approx 1024$, we compare it to the theoretical maps at $n_{\text{side}} = 256$ for computational efficiency. In Section VI B 1 we find that out results are not significantly affected by this choice.

Because of the high density and close proximity of the centre of our own galaxy, one would expect that a DM annihilation or decay signal would be dominated by this region. However, the constraint or detection one would obtain from studying this region would be sensitive to the modelling of the Milky Way density profile, and one would have to ensure that such a signal could not arise due to potentially incorrect modelling of the galactic diffuse or isotropic components, or through processes not captured by these models, such as an unresolved population of millisecond pulsars. This is a complicated yet feasible task \([\text{see e.g. 17}]\), but is beyond the scope of this work; here we wish to produce constraints on DM annihilation and decay that are independent of the GCE so we simply mask the galactic plane, with the aim that any constraint or detection is driven by the density fields inferred in Section IV A. We therefore mask the region with galactic latitude $|\lambda| < 30^\circ$.

IV. METHODS

In this section we detail how we construct the full-sky templates for DM annihilation and decay and how these are compared to the gamma-ray data to constrain the annihilation cross section and decay rate. In Section IV A we describe the constrained simulations used to generate these templates and in Section IV B we describe how the $J$ and $D$ factors are computed from the DM particles in these simulations. These templates are combined with those from Section IV C that describe non-DM contributions to the gamma-ray sky, and we compare these to the data using the likelihood model in Section IV D.

A. Bayesian large-scale structure inference

To compute the $J$ and $D$ factors, we use the results of applying the BORG algorithm \([\text{see e.g. 59–63}]\) to the 2M++ galaxy compilation \([63, 69]\). The BORG algorithm applies a Bayesian forward model for the observed number densities of galaxies to infer both the present-day three-dimensional density field and the corresponding ICs. The algorithm incorporates a physical model for gravitational structure formation and marginalises over galaxy bias parameters.

Specifically, we use a set of DM-only constrained simulations, dubbed CS1BORG \([55–57]\). These start by taking 101 sets of $z = 69$ ICs from across the BORG chain, covering a box length of $677.7\,\text{h}^{-1}\,\text{Mpc}$ with $256^3$ voxels. Inside a sphere of radius $155\,\text{h}^{-1}\,\text{Mpc}$ centred on the Milky Way, these ICs are augmented with white noise to a resolution of $2048^3$, resulting in a particle mass of $4.38 \times 10^9\,\text{M}_\odot$. These simulations are then run to $z = 0$ using RAMSES \([70]\), giving 101 $N$-body realisations of the local DM distribution. Only the central, high-resolution region is refined; the rest of the BORG box is retained only to include the effect of longer-wavelength modes. Both BORG and CS1BORG adopt the cosmology $T_{\text{CMB}} = 2.728\,\text{K}$, $\Omega_m = 0.307$, $\Omega_A = 0.693$, $\Omega_b = 0.04825$, $H_0 = 70.5\,\text{km}\,\text{s}^{-1}\,\text{Mpc}^{-1}$, $\sigma_8 = 0.8288$ and $n_s = 0.9611$. Our results are therefore conditioned on this cosmology; a study of the cosmology dependence of our results is beyond the scope of this work.

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1. https://fermi.gsfc.nasa.gov/ssc/data/analysis/software/
2. http://healpix.sf.net
To improve the effective resolution of our calculation in high-density regions we run a halo finder on the CSiBORG particles and use analytic formulae for a specified density profile. Specifically, we use the watershed halo finder PH EW [71] which runs on the fly as a patch to RAMSES. This splits the negative density field into basins that share a common local minimum via steepest descent. These basins are then combined according to user-defined thresholds on the density of saddle points between basins to merge low-mass subhalos into their parents; we use the standard threshold value from Bleuler et al. [71] of 200\(\rho_c\). The halo catalogues are publicly available for the full CSiBORG suite [58].

### B. Computing the \(J\) and \(D\) factors

The halo finding allows us to split CSiBORG particles into two types: those that belong to halos and those that do not. Since the \(J\) factor depends on the square of the density, it is more sensitive to the small-scale matter distribution, and thus we must treat halos separately from the background density field in order to account for structures below the resolution of the CSiBORG simulations. The \(D\) factor is less sensitive to these small scales, and thus we treat all particles equally in this case. \(D\) is computed using the procedure outlined in Section IV B 1, where we use all particles. We compute \(J\) as in Section IV B 1, but only considering nonhalo particles, and add this to the contribution from halos, which are treated as in Section IV B 2. We plot the resulting ensemble mean \(J\) and \(D\) factor maps in Fig. 1.

For the \(J\) factor maps, we generate the templates on a higher resolution HEALPix grid than that on which we perform the inference (\(\text{nside} = 2048\) instead of \(\text{nside} = 256\)) and subsequently degrade them. Because of the nonlinear dependence of \(J\) on \(\rho\), this allows for a more faithful representation of the density field than if \(J\) was initially calculated at \(\text{nside} = 256\), which is especially important for the regions of the sky corresponding to halos produced in CSiBORG. We concluded the HEALPix resolution of \(\text{nside} = 2048\) was sufficient by comparing the total \(J\) factor to those with increasing HEALPix resolution (\(\text{nside} = 4096, 8192\)) since the change in total \(J\) factor was at the subpercent level with increasing \(\text{nside}\).

Regarding the \(D\) factor calculation, we directly calculate the line-of-sight integral of the density field within a given HEALPix pixel at the selected resolution. The convergence of the total all-sky \(D\) factor with this procedure is, of course, present for all considered resolutions, since it should be simply proportional to the total mass within the CSiBORG volume.

## 1. Smoothed density field

We wish to determine the density of DM particles that do not belong to a halo on a regular Cartesian grid with \(N_{\text{grid}} = 1024\) grid points per side. To do this, we adopt a procedure based on smooth particle hydrodynamics (SPH) [72] as described in [73] and outlined below. Using the SPH algorithm over, e.g., a cloud-in-cell (CIC) approach allows us to better capture the peaks of the matter density field, since the SPH kernel will adapt to the local density of matter, in contrast to the CIC approach which has a fixed kernel corresponding to a trilinear interpolation scheme. We compare the results of using a SPH kernel to a CIC algorithm in Section VI B 1.

First, we determine the number of particles, \(N_p\), within the cell corresponding to each grid point \((i, j, k)\). We then define

\[
N_X = \max(N_p, N_{\text{SPH}}),
\]

where \(N_{\text{SPH}} = 32\). The choice for this number of neighbours is partly motivated by the typical number of edges linking a node to its neighbours in a Delaunay tessellation. That number is approximately 16 for a Euclidean three-dimensional vector space [74, 75]. We pick a value twice as big as we intend the filter to have a larger reach than the first layer of neighbours. We then find the mass associated with this grid point by considering the nearest \(N_X\) particles to be

\[
\tilde{m}_{ijk} = \frac{1}{R_{ijk}^3} \sum_{l=1}^{N_X-1} m_l W_l \mathcal{S} \left( \frac{d_l}{R_{ijk}} \right),
\]

where \(R_{ijk}\) is half the distance to the farthest of the \(N_X\) particles from the grid point, \(m_l\) is the mass of particle \(l\), which is at a distance \(d_l\) from the grid point, \(W_l\) is the weight for particle \(l\),

\[
W_l = \left( \sum_{ijk} \frac{R_{ijk}^3}{R_{ijk}^3} \mathcal{S} \left( \frac{d_l}{R_{ijk}} \right) \right)^{-1},
\]

and the interpolating function, \(\mathcal{S}\), is chosen to be

\[
\mathcal{S}(x) = \begin{cases} 
1 - \frac{3}{2}x + \frac{3}{2}x^3, & 0 \leq x < 1 \\
\frac{1}{4}(2 - x)^3, & 1 \leq x < 2 \\
0, & \text{otherwise}.
\end{cases}
\]

If the spacing between grid points is \(\Delta r\), then the density assigned to each site is

\[
\hat{\rho}_{ijk} = \frac{\tilde{m}_{ijk}}{\Delta r^3}.
\]

To compute the \(J\) and \(D\) factors, we compute Eqs. 5 and 6, respectively, along the line of sight corresponding to each HEALPix pixel at the chosen resolution. We integrate up to the edge of the simulated volume and perform trilinear interpolation of the density field onto
In this section we review how the completeness used in the BORG inference of the initial conditions [63, 69] (faded region near the galactic plane masked out).

2. $J$ factor from halos

To include the contribution from particles inside halos, we use a custom extension of the CLUMPY package\(^3\) [76–78]. In this section we review how the $J$ factor is calculated in this package, as well as our assumptions for the halo density profiles.

We assume that all halos are spherically symmetric and that the total density profile, $\rho_{\text{tot}}$, can be described by a simple analytic form. For our fiducial case, we consider the three-parameter family of profiles [79, 80]

$$
\rho_{\alpha\beta\gamma} (r) = \frac{2^{\frac{\alpha - 5}{\beta}} \rho_s}{(r/r_s)^\gamma (1 + (r/r_s)^\alpha)^{\frac{2-\alpha}{\beta}}},
$$

where $r_s$ is the scale radius, $\rho_s$ is the density at $r_s$, and $\alpha$ describes the sharpness of the transition between the inner (\(\gamma\)) and outer (\(\beta\)) logarithmic slopes. For a NFW profile, $\alpha = 1$, $\beta = 3$, $\gamma = 1$. In Section VI B 1, we also consider the Einasto profile [54, 81]

$$
\rho_{\text{Einasto}} (r) = \rho_{-2} \exp \left( -\frac{2}{\alpha} \left( \left(\frac{r}{r_{-2}}\right)^{\alpha} - 1 \right) \right).
$$

In the above, $r_{-2}$ represents the radius at which the logarithmic slope of the profile equals $-2$, while $\rho_{-2} \equiv \rho (r_{-2})$. For both profiles we calculate the parameters defining the profile from the total halo mass and corresponding concentration. The mass-concentration relation we use is shown in Eq. 25. We also note that the parameter $\alpha$ of the Einasto profile is determined as a function of virial peak height [82]. Given that all of our halos are at very low redshift ($z \lesssim 0.05$) and almost all have masses within $M \in [10^{12}, 10^{15}] M_\odot$, the parametric relation is quite accurate. We use the COLOSSUS package throughout [83]. We note that, since the Einasto parameters are fitted to halos produced in DM-only simulations, this captures only the uncertainty in N-body modelling. Baryons induce a potentially larger effect, which is, however, harder to model reliably. We discuss this further in Section VI B 2.

One would also expect that a halo contains a large number of “clumps” or subhalos, such that the true smooth component of the density profile is [84]

$$
\rho_{\text{sm}} (r) = \rho_{\text{tot}} (r) - \langle \rho_{\text{subs}} (r) \rangle,
$$

where $\langle \rho_{\text{subs}} (r) \rangle$ gives the average contribution from the substructure. If clump $i$ has density profile $\rho_{cl}^i$, then it contributes to the total $J$ factor value within a HEALPix pixel, $p$, as

$$
J_p = \int_{p, \Delta_{\text{halo}}} \rho_{\text{sm}} (s, \Omega) + \sum_i \rho_{cl}^i (s, \Omega) \, dsd\Omega
= J_{\text{sm, } p} + J_{\text{subs, } p} + J_{\text{cross, } p},
$$

where

$$
J_{\text{sm, } p} = \int_{p, \Delta_{\text{halo}}} \rho_{\text{sm}}^2 (s, \Omega) \, dsd\Omega,
$$

$$
J_{\text{subs, } p} = \int_{p, \Delta_{\text{halo}}} \left( \sum_i \rho_{cl}^i (s, \Omega) \right)^2 \, dsd\Omega,
$$

$$
J_{\text{cross, } p} = 2 \int_{p, \Delta_{\text{halo}}} \rho_{\text{sm}} (s, \Omega) \left( \sum_i \rho_{cl}^i (s, \Omega) \right) \, dsd\Omega,
$$

\(^3\) https://clumpy.gitlab.io/CLUMPY
and \( \Delta_{\text{halo}} \) represents the intersection of the halo volume with the cone spanned by the pixel \( p \). Our task is therefore to determine the distribution of clumps for a given halo and to calculate these integrals. Here, we readily use the solution provided by the CLUMPY package and describe it briefly below.

Assuming that a given halo has \( N_{\text{tot}} \) independent clumps, we factorise the distribution for the number of clumps with some mass \( M \), concentration \( c \), in some region \( dV = d^3r \) as [77, 85]

\[
\frac{dN}{dV dM dc} = N_{\text{tot}} \frac{dP_V(r)}{dV} \frac{dP_M(M)}{dM} \frac{dP_c(M,c)}{dc}.
\]

(21)

Since the clumps form before the host halos within ΛCDM, their spatial distribution will follow the host DM density profile. This has been shown to be a good assumption in simulations of galaxy-sized halos [54, 86]. Given the self-similar nature of collapse of ΛCDM halos, we extrapolate this conclusion to halos from our CSiBORG ensemble. We assume that the distribution of masses is a power law

\[
\frac{dP_M}{dM} \propto M^{-\alpha_M},
\]

(22)
in the range \( M \in [10^{-6}M_\odot, 10^{-2}M_\odot] \) for a halo of mass \( M_h \), where \( \alpha_M = 1.9 \) (see Sec. 2.3 of [87] and references therein). Again, the values are motivated by numerical simulations of Milky Way sized halos, which we extrapolate to bigger halos present in our forward model.

Besides modelling the uncertainty due to the spatial and mass distribution of substructure, the CLUMPY package also allows us to include the uncertainty in the mass-concentration relation. For the substructure component, we consider two cases for the concentration distribution. In the first case, we assume that the concentration of all substructure halos is a deterministic function of the mass

\[
\frac{dP_c}{dc} = \delta(c - \tilde{c}(M)),
\]

(23)

where we define \( \delta \) as the Dirac-delta distribution. The second possibility that we follow is that the concentration is log normally distributed about this mean

\[
\frac{dP_c}{dc} = \frac{1}{\sqrt{2\pi}\sigma_c(M)} \exp\left(-\frac{(\log c - \log(\tilde{c}(M)))^2}{2\sigma_c^2}\right).
\]

(24)

Motivated by [51], the substructure halos are assumed to have the following mass-concentration relation [88]

\[
\tilde{c}(M) = \sum_{j=0}^{5} C_j \left[ \ln \left( \frac{M}{\rho_c h^{-1} M_\odot} \right) \right]^j,
\]

(25)

where \( C_j = (37.5153, -1.5093, 1.636 \times 10^{-2}, 3.66 \times 10^{-4}, -2.8927 \times 10^{-5}, 5.32 \times 10^{-7}) \), with \( \sigma_c = 0.0 \); i.e., we assume all substructure halos of the same mass have the same concentration. In Section VI B 3 we consider a nonzero value \( \sigma_c = 0.2 \) in Eq. 24, as motivated by [89–91] as a comparison.

Given that we do not resolve substructures of all the halos present with our simulation, we assume the resulting distribution for \( J_{\text{sm}} \) and \( J_{\text{subs}} \) to be a Gaussian. Therefore, we only need to find the mean and variance of these contributions in each HEALPix pixel. We define the 1-clump luminosity to be

\[
L(M,c) \equiv \int_{V_{\text{subhalo}}} \rho_{\text{cl}}^2(r; M,c) d^3r,
\]

(26)

and its moments as

\[
\langle L^n \rangle \equiv \int_{M_1}^{M_2} \frac{dP_M}{dM} \int \frac{dP_c}{dc} L^n dc dM,
\]

(27)

for a given mass range of clumps \([M_1, M_2] \), while the
mean contribution of \( J_{\text{subs},p} \) from this volume is
\[
\langle J_{\text{subs},p} \rangle = N_{\text{tot}} \int_{\rho, \nu_h} dP_V \ell^{-2} dV \langle L \rangle,
\]
with \( N_{\text{tot}} \) representing the total number of clumps within the selected mass range \([M_1, M_2]\). Note that we assume that the clumps are nonoverlapping, such that the cross terms in Eq. 19 can be neglected. For more details on how these quantities are defined we refer the reader to the CLUMPY related publications [76–78].

Note that since we are assuming unresolved substructures for our CSiBORG extragalactic halos, we are integrating over the total subhalo volume \( V_{\text{subhalo}} \) for the subhalo luminosity. Furthermore, since there can be many subhalos present within the line of sight determined by the given HEALPix pixel we are also accounting for the span of the host halo along this line of sight through \( V_{\ell} \equiv [\ell_{\text{min}}, \ell_{\text{max}}] \), with \( \ell_{\text{min}} \) and \( \ell_{\text{max}} \) being the closest and farthest points of the host halo along this line of sight. Since these integrals do not have a closed form for general DM profiles, we evaluate all numerically.

Given that the mean of some power of the distance from the observer, \( \ell \), to a clump which falls inside the HEALPix pixel \( p \) is
\[
\langle \ell^n_\rho \rangle = \int_{\rho, \nu_h} \ell^{n+2} dP_V d\ell d\Omega,
\]
we can write the variance on \( J_{\text{subs}} \) as
\[
\sigma^2_{J_{\text{subs}}} = \langle L^2 \rangle \langle \ell^{-4}_p \rangle - \langle L^2 \rangle \langle \ell^{-2}_p \rangle^2,
\]
since \( L \) and \( \ell \) are independent. For the cross term, \( J_{\text{cross},p} \), we use that its mean is
\[
\langle J_{\text{cross},p} \rangle = 2 \int_{\rho, \nu_h} \rho_{\text{sm}} \langle \rho_{\text{subs}} \rangle d\ell d\Omega,
\]
while its variance can be computed as
\[
\sigma^2_{J_{\text{cross}}} = 4 \int_{\rho, \nu_h} \rho_{\text{sm}}^2 \langle l, \Omega \rangle \sigma^2_{\text{subs}}(l, \Omega) d\ell d\Omega,
\]
with
\[
\sigma^2_{\text{subs}} \equiv \sigma^2_{\text{subs}}(l, \Omega) = \langle \rho_{\text{subs}}^2(l, \Omega) \rangle - \langle \rho_{\text{subs}}(l, \Omega) \rangle^2,
\]
and
\[
\langle \rho_{\text{subs}}(l, \Delta \Omega) \rangle = \int_{\Delta \Omega, \nu_h} \frac{dP_V}{d\Delta \Omega} \int_{\nu_h} \frac{dP_M}{dM} \int_{\nu_h} \frac{dP_c}{dc} \times \rho_{\text{subs}}(\ell, \Omega; M, c(M)) d\nu_h dM dc,
\]
with \( \nu_h \) and \( \nu_h(M) \) representing the mass and corresponding concentration ranges for the subhalo distribution respectively.

We decide to include only \( \sigma_{J_{\text{subs}},p} \) as it is the dominant source of uncertainty. This can intuitively be understood from the Eq. 32. We note that the integrand is negligible both in the outskirts of the host halo, since \( \rho_{\text{sm}} \sim 0 \) and \( \sigma^2_{\text{subs}} \) remains finite, and in the very centre of the host halo, since \( \sigma_{\text{subs},p} \sim 0 \). Therefore, \( J_{\text{cross},p} \) contributes only at a very limited range of scales. Furthermore, for our halos, there is a clear hierarchy between the cross and sub terms \( J_{\text{cross},p} \lesssim 0.01 - 0.1 J_{\text{subs},p} \); therefore, we focus only on the \( \sigma_{J_{\text{subs}}} \) as the dominant source of uncertainty of the \( J \) factor due to substructure.

We hence write that the distribution followed by the \( J \) factor for a given halo in a given pixel is given by a Gaussian with mean
\[
\langle J_p \rangle = J_{\text{sm},p} + \langle J_{\text{subs},p} \rangle + \langle J_{\text{cross},p} \rangle,
\]
and variance
\[
\sigma^2_p = \sigma^2_{J_{\text{subs}}} + \sigma^2_{J_{\text{cross}}} \approx \sigma^2_{J_{\text{subs}}}.
\]
We calculate these quantities for all CSiBORG realisations. To distinguish between these, we introduce a second index, \( j \), to label the simulation, i.e. \( \langle J_p \rangle \) is \( \langle J_p \rangle_j \) for CSiBORG simulation \( j \), and likewise \( \sigma_p \) is \( \sigma_p \) for the same simulation.

In Fig. 2, we show the result of the model for a typical halo within CSiBORG with a mass of \( M_h \approx 5 \times 10^{13} M_\odot \). We see that the effects from the term in \( \sigma^2_{\text{subs}} \) cannot be neglected, especially in the outskirts of the halo. In the very centre, where the structure of the halo is dominated by the smooth component, the fluctuations in the \( J \) factor due to the substructure are negligible, amounting to only few percent, while in the outer parts these fluctuations become more important. This is an expected result given that the boost in \( J \) factor due to substructure becomes more important in the outer edges, where the smooth component, \( J_{\text{sm},p} \), is subdominant with respect to the substructure \( J \) factor, \( J_{\text{subs},p} \). This behaviour is identical for an Einasto profile.

Besides this, we also include the contributions of subsubclumps to the \( J \) factor of halos, using one additional level of substructure, which is the default setting of the CLUMPY code. Because of the increased computational cost, we considered a two-level substructure contribution for our halos for only one CSiBORG realisation. Including additional substructure levels will result in an overall change in the \( J \) factor of less than \( \sim 5\% \) [77], and ignoring such levels will make our constraints on \( \langle \sigma V \rangle \) conservative since this will systematically underestimate the \( J \) factor.

In conclusion, to obtain the total all-sky \( J \) factor, we combine the line-of-sight calculation for the density field obtained from particles outside of halos detected within CSiBORG realisations with the component coming from the halo particles of the CSiBORG by treating the halos as presented in this section, utilising the CLUMPY code. This final template is used in the inference pipeline. We discuss the numerical convergence of these calculations in Section VIB.
include a filtered residual map, which is smoothed to a few or large-scale diffuse structures, since the templates in our galaxy. We note that these templates are not de- 

each energy bin, which controls the total emission from unresolved extragalactic sources, residual cosmic rays and extragalactic diffuse sources. It consists or a spatially constant map, with a spectral shape given by the Fermi Isotropic Spectral Template$^4$, but with an overall normalisation $A^{\text{iso}}$, which we infer separately for each energy bin, $i$. 

The isotropic component is designed to capture emission from unresolved extragalactic sources, residual cosmic rays and extragalactic diffuse sources. It consists of a spatially constant map, with a spectral shape given by the Fermi Isotropic Spectral Template$^4$, but with an overall normalisation $A^{\text{iso}}$, which we infer separately for each energy bin, $i$. 

For the galactic component, we use the spatial models described in [92], which are developed using spectral line surveys of HI and CO and infrared tracers of dust column density and a model of inverse Compton emission. These spatial templates describe the relative change in flux across different parts of the sky. We keep these fixed during our analysis and infer the normalisation, $A^{\text{gal}}$, in each energy bin, $i$, which controls the total emission from our galaxy. We note that these templates are not designed to be used for analyses which aim to fit medium or large-scale diffuse structures, since the templates include a filtered residual map, which is smoothed to a few degrees. Any large-scale component not explicitly modelled when generating the templates will be absorbed by this residual and would be undetectable in our analysis. However, as can be seen in Fig. 1, the $J$ and $D$ factor maps are dominated by small-scale features due to massive structures, and thus we are in the regime for which these templates are valid. 

Finally, we produce a single template containing all point and extended sources from the models provided in the LAT 12-year Source Catalog (4FGL-DR3)$^5$. We introduce a free scaling parameter, $A^{\text{psc}}$, for each energy bin which will be inferred. Since we find no cross-correlation between the point-source template and our own $J$ and $D$ factor templates, having a separate normalisation for the point-source template will not introduce new degeneracies (see Section VI B 4).

### D. Likelihood model

Instead of directly constraining the DM annihilation or decay parameters, we split the inference into two parts. First, we assume that there is a contribution to the gamma-ray sky which is proportional to the $J$ or $D$ factor, i.e. for a given CSiBORG simulation $j$, the flux in energy bin $i$ and pixel $p$, $\Phi_{ipj}$, has terms $\Phi_{ipj} \sim \left(A_i^J J_{pj}^{ip} + A_i^D D_{pj}^{ip}\right) A_p \Delta E_i$, (37)

where $\Delta E_i$ is the width of the bin, $A_p$ is the area of the pixel in steradians and $J_0$ and $D_0$ set the units. We choose $J_0 = 10^{13}$ GeV$^2$ cm$^{-5}$ and $D_0 = 10^{13}$ GeV cm$^{-2}$. We fit for the total flux of such a contribution in each energy bin to obtain a spectrum. In the second half of the inference, we fit this spectrum to a series of models (including DM annihilation and decay) in attempt to determine the origin of such a signal.

This method has two main advantages. First, we can consider each energy bin and CSiBORG simulation separately in the first half of the inference. Although this involves initially computing 909 MCMC chains (one for each energy bin and for each CSiBORG simulation), since the problem is embarrassingly parallelisable and because we only need to infer four or five parameters for each chain ($A_i^J$ and/or $A_i^D$, $A_i^{\text{iso}}$, $A_i^{\text{gal}}$, $A_i^{\text{psc}}$) compared to $\sim 30$ if we combined the energy bins, we find that this approach is computationally more efficient. Second, by remaining agnostic to the origin of $A_i^J$ or $A_i^D$ until the second step, we are able to more easily determine which energy bins drive our constraints. Hence, it becomes simpler to compare different models since we do not need to rerun the map-level inference every time that we wish to change the DM particle mass or decay channel (which is more computationally expensive).

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$^4$ [https://fermi.gsfc.nasa.gov/ssc/data/access/lat/BackgroundModels.html](https://fermi.gsfc.nasa.gov/ssc/data/access/lat/BackgroundModels.html)

$^5$ [https://heasarc.gsfc.nasa.gov/W3Browse/fermi/fermilpsc.html](https://heasarc.gsfc.nasa.gov/W3Browse/fermi/fermilpsc.html)

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FIG. 3. Mollweide projection in galactic coordinates of the template predictions across the full energy range considered in this work. For visualisation, we sum the isotropic, galactic diffuse, and point source templates assuming each has a unit amplitude. In our inference we simultaneously infer the normalisation of each of the three components, with a different amplitude for each energy bin, and the contribution proportional to the $J$ or $D$ factor.

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### C. Non-DM templates

To constrain the parameters describing DM annihilation or decay, one also needs to take into account other sources of gamma rays. We consider a model with three contributions: our own galaxy (gal), an isotropic background (iso) and point sources (psc). We produce separate templates, $\{T^t_i(\hat{r})\}$, $t \in \{\text{iso, gal, psc}\}$ for each energy bin, $i$, and assign each template a different normalisation, which we infer from the data. The sum of the three templates is plotted in Fig. 3. 

The isotropic component is designed to capture emission from unresolved extragalactic sources, residual cosmic rays and extragalactic diffuse sources. It consists of a spatially constant map, with a spectral shape given by the Fermi Isotropic Spectral Template$^4$, but with an overall normalisation $A^{\text{iso}}$, which we infer separately for each energy bin, $i$. 

For the galactic component, we use the spatial models described in [92], which are developed using spectral line surveys of HI and CO and infrared tracers of dust column density and a model of inverse Compton emission. These spatial templates describe the relative change in flux across different parts of the sky. We keep these fixed during our analysis and infer the normalisation, $A^{\text{gal}}$, in each energy bin, $i$, which controls the total emission from our galaxy. We note that these templates are not designed to be used for analyses which aim to fit medium or large-scale diffuse structures, since the templates include a filtered residual map, which is smoothed to a few degrees. Any large-scale component not explicitly modelled when generating the templates will be absorbed by this residual and would be undetectable in our analysis. However, as can be seen in Fig. 1, the $J$ and $D$ factor maps are dominated by small-scale features due to massive structures, and thus we are in the regime for which these templates are valid.

Finally, we produce a single template containing all point and extended sources from the models provided in the LAT 12-year Source Catalog (4FGL-DR3)$^5$. We introduce a free scaling parameter, $A^{\text{psc}}$, for each energy bin which will be inferred. Since we find no cross-correlation between the point-source template and our own $J$ and $D$ factor templates, having a separate normalisation for the point-source template will not introduce new degeneracies (see Section VI B 4).

### D. Likelihood model

Instead of directly constraining the DM annihilation or decay parameters, we split the inference into two parts. First, we assume that there is a contribution to the gamma-ray sky which is proportional to the $J$ or $D$ factor, i.e. for a given CSiBORG simulation $j$, the flux in energy bin $i$ and pixel $p$, $\Phi_{ipj}$, has terms $\Phi_{ipj} \sim \left(A_i^J J_{pj}^{ip} + A_i^D D_{pj}^{ip}\right) A_p \Delta E_i$, (37)

where $\Delta E_i$ is the width of the bin, $A_p$ is the area of the pixel in steradians and $J_0$ and $D_0$ set the units. We choose $J_0 = 10^{13}$ GeV$^2$ cm$^{-5}$ and $D_0 = 10^{13}$ GeV cm$^{-2}$. We fit for the total flux of such a contribution in each energy bin to obtain a spectrum. In the second half of the inference, we fit this spectrum to a series of models (including DM annihilation and decay) in attempt to determine the origin of such a signal.

This method has two main advantages. First, we can consider each energy bin and CSiBORG simulation separately in the first half of the inference. Although this involves initially computing 909 MCMC chains (one for each energy bin and for each CSiBORG simulation), since the problem is embarrassingly parallelisable and because we only need to infer four or five parameters for each chain ($A_i^J$ and/or $A_i^D$, $A_i^{\text{iso}}$, $A_i^{\text{gal}}$, $A_i^{\text{psc}}$) compared to $\sim 30$ if we combined the energy bins, we find that this approach is computationally more efficient. Second, by remaining agnostic to the origin of $A_i^J$ or $A_i^D$ until the second step, we are able to more easily determine which energy bins drive our constraints. Hence, it becomes simpler to compare different models since we do not need to rerun the map-level inference every time that we wish to change the DM particle mass or decay channel (which is more computationally expensive).
1. Inferring the Template

We assume that photon counts in energy bin $i$ from the $J$ and $D$ factor contributions, as well as each of the contributions described in Section IV C is Poisson distributed. The variation of the mean of the latter with sky position, $\hat{r}$, and energy is described by the known templates $\{T_i^t(\hat{r})\}$, where $t$ labels the templates. For pixel $p$ and CSiBORG realisation $j$, we then define

$$J_{ipj} \equiv \frac{J_{pj}}{J_0} \Delta E_i, \quad D_{ipj} \equiv \frac{D_{pj}}{D_0} \Delta E_i, \quad T_{ip}^t = \int_p T_{i}^t(\hat{r}) \, d\Omega,$$

such that the mean number counts in pixel $p$ and energy bin $i$ is predicted to be

$$\lambda_{ipj} = F_{ip} \times \left( A_i^J J_{ipj} + A_i^D D_{ipj} + \sum_t A_i^T_{ip} T_{ip}^t \right),$$

and we have multiplied our templates by the Fermi exposure, $F_{ip}$, which describes the angular variation of the sensitivity of Fermi. This step is performed using the Fermi Tools, where we also convolve all templates with the point spread function. The likelihood of observing $n_{ipj}$ counts in pixel $p$ and energy bin $i$ given the mean $\lambda_{ipj}$ is

$$L(n_{ipj}|\lambda_{ipj}) = \frac{\lambda_{ipj}^{n_{ipj}} \exp(-\lambda_{ipj})}{n_{ipj}!}. \quad (40)$$

As discussed in Section IV B, we do not know the exact DM distribution for a given CSiBORG simulation due to the unresolved substructure in halos, although we did not include this uncertainty in Eq. 40. We model the uncertainty on the substructure contribution to the $J$ factor as a truncated Gaussian. This choice allows us to marginalise analytically over the substructure uncertainty, such that the conditional probability for a given $\lambda_{ipj}$ given the mean parameters is

$$L(\lambda_{ipj}|A_i^J, A_i^D, \{A_i^t\}, j) = \frac{1}{\sigma_{ipj}} \sqrt{\frac{2}{\pi}} \times \left( 1 + \text{erf} \left( \frac{\mu_{ipj}}{\sigma_{ipj} \sqrt{2}} \right) \right) - 1 \exp \left( -\frac{(\lambda_{ipj} - \mu_{ipj})^2}{2 \sigma_{ipj}^2} \right), \quad (41)$$

for $\lambda_{ipj} > 0$, and zero otherwise, where

$$\mu_{ipj} = F_{ip} \times \left( A_i^J \langle J_{ipj} \rangle + A_i^D \langle D_{ipj} \rangle + \sum_t A_i^T_{ip} T_{ip}^t \right),$$

$$\sigma_{ipj} = A_i^J \frac{\sigma_{pj}}{J_0} \Delta E_i.$$

The likelihood for observing $n_{ipj}$ counts in pixel $p$ and energy bin $i$ is then

$$L(n_{ipj}|A_i^J, A_i^D, \{A_i^t\}, j) = \int L(n_{ipj}|\lambda_{ipj}) \, L(\lambda_{ipj}|A_i^J, A_i^D, \{A_i^t\}, j) \, d\lambda_{ipj}$$

$$= \sqrt{\frac{2\pi \sigma_{ipj}^2}{n_{ipj}}} \exp \left( -\frac{\mu_{ipj}^2}{2 \sigma_{ipj}^2} \right) \left( 1 + \text{erf} \left( \frac{\mu_{ipj}}{\sigma_{ipj} \sqrt{2}} \right) \right) \times \left( \Gamma \left( \frac{n_{ipj} + 1}{2} \right) 1 F_1 \left( \frac{n_{ipj} + 1}{2}, \frac{1}{2} \left( \frac{\mu_{ipj} - \sigma_{ipj}^2}{\sigma_{ipj}^2} \right) \right) \right.$$

$$\times \left. \sqrt{\frac{1}{2}} \frac{(\mu_{ipj} - \sigma_{ipj}^2)}{\sigma_{ipj}} \Gamma \left( \frac{n_{ipj} + 1}{2} \right) \right), \quad (44)$$

where $1 F_1$ is the confluent hypergeometric function of the first kind. We describe how we implement this likelihood numerically in Appendix A.

Assuming that each pixel is independent, the likelihood for the observed data in energy bin $i$, $D_i$, is

$$L(D_i|A_i^J, A_i^D, \{A_i^t\}, j) = \prod_p L(n_{ipj}|A_i^J, A_i^D, \{A_i^t\}, j).$$

Using the priors, $P$, given in Table I, we apply Bayes’ identity

$$L(A_i^J, A_i^D, \{A_i^t\}, j|D_i) = \frac{L(D_i|A_i^J, A_i^D, \{A_i^t\}, j) \, P(A_i^J) \, P(A_i^D) \, P(\{A_i^t\}) \, P(j)}{Z(D_i)}, \quad (46)$$

where

$$P(\{A_i^t\}) = \prod_t P(A_i^t), \quad (47)$$

to obtain the posterior, $P(A_i^J, A_i^D, \{A_i^t\}, j|D_i)$, where $Z(D_i)$ is the evidence and where we consider each energy bin and CSiBORG simulation separately. We use the emcee sampler [93] and terminate the chain when the estimate of the autocorrelation length changes by less than 1 percent per iteration and the chain is at least 100 autocorrelation lengths long in all of the parameters.

We apply a Monte Carlo estimate to the likelihood of the CSiBORG samples, such that the one-dimensional posterior for the amplitude $A_i^X$ ($X$ being $J$ or $D$) is

$$L(A_i^X|D_i) = \frac{1}{N_{\text{sim}}} \sum_j d\{A_i^j\} dA_i^X \, L(A_i^J, A_i^D, \{A_i^t\}, j|D_i), \quad (48)$$

where $Y = J$ if $X = D$ and vice versa. In practice, we compute the average over CSiBORG realisations by first
fitting the one-dimensional posteriors \( L(A_i^X | D_i, j) \) with a spline using the GetDist package [94] and then computing the mean of the resulting functions. This is equivalent to concatenating the Markov Chains if each chain had the same length.

2. Constraining DM parameters

We now have a posterior, \( L(A_i^1 | D_i) \), describing the gamma-ray spectrum from sources that have the same spatial distribution as the \( J \) factor. We wish to fit this spectrum to a model, \( f_i(\theta) \), for these sources and infer the model parameters \( \theta \). We assume that our model is deterministic, such that

\[
 L(A_i^1 | \theta) = \delta (A_i^1 - f_i(\theta)),
\]

and therefore we obtain the likelihood for the observed gamma-ray sky by incorporating Eq. 48,

\[
 L(D_i | \theta) = \int dA_i^1 L(D_i | A_i^1) L(A_i^1 | \theta) = \int dA_i^1 \frac{L(A_i^1 | D_i) Z(D_i)}{P(A_i^1)} \delta (A_i^1 - f_i(\theta)).
\]

We assume that all energy bins are independent such that the likelihood of \( \theta \) given the full dataset \( D \) is

\[
 L(D | \theta) = \prod_i L(D_i | \theta),
\]

and so with a final application of Bayes’ identity we obtain the posterior for our model parameters

\[
 L(\theta | D) = \frac{L(D | \theta) P(\theta)}{Z(D)}.
\]

If \( f_i(\theta) \) comprises exclusively of DM annihilation, then, at fixed DM mass and annihilation channel, the transformation from the posterior distribution of \( A_i^1 \) to that of \( \langle \sigma v \rangle \) is trivial. For more complicated models (i.e., where \( \theta \) consists of more than one parameter), we again calculate the posterior on \( \theta \) using the emcee package.

For DM annihilation and decay, prompt production, decays, hadronisation and radiative processes associated with the resulting standard model products produce a variety of stable species, including gamma rays. For a given channel, one must know the energy spectrum of the intermediate standard model particles and the resulting branching ratios and energies of the subsequently produced particles. One then has, for each channel, a model for the spectrum of gamma rays as a function of DM particle mass and annihilation cross-section or decay rate.

In this work we utilise the pre-computed spectra provided by the Fermi collaboration\(^6\) which are calculated as described by Jeltema and Profumo [95]. Since we are considering non-relativistic s-wave annihilation in this work, one can view the annihilation of two DM particles of mass \( m_\chi \), as equivalent to the decay of a single particle of mass \( 2m_\chi \) [96]. Hence, we obtain the spectrum for decay from the tabulated annihilation spectra by evaluating these at half the relevant particle mass. Furthermore, for kinematic reasons, if we produce two standard model particles, each of rest mass \( m_A \), then for decay we enforce the DM particle mass to obey \( m_\chi > 2m_A \), whereas for annihilation this is \( m_\chi > m_A \).

## V. RESULTS

In Fig. 4, we show the corner plot for the first stage of our inference, where we infer \( A_i^1 \) and \( A_i^D \) simultaneously. We emphasise that we fit a different \( A_i^1 \) and \( A_i^D \) to each energy bin, \( i \), and CS1BORG simulation. In this example, we consider simulation number 7444 (as given in [58]) and the energy range \( 30 - 50 \text{ GeV} \) (energy bin 9). We see that the parameters corresponding to the isotropic, galactic and point-source contributions are all approximately unity, as one would expect. For this energy bin and CS1BORG simulation we see that there is no evidence for a contribution to the gamma-ray flux proportional to either the \( J \) or the \( D \) factor. We note that \( A_i^1 \) and \( A_i^D \) are highly degenerate, such that a large value of \( A_i^1 \) corresponds to a small \( A_i^D \). For our fiducial analysis, we therefore choose to set one of these parameters equal to zero at a time; i.e. the inference to place constraints on \( \langle \sigma v \rangle \) will assume \( \forall i \ A_i^D = 0 \) and for \( \Gamma \) we assume \( \forall i \ A_i^1 = 0 \). This will make our constraints conservative (see Section VI A 1).

We note that \( A_i^{iso} \) is strongly degenerate with \( A_i^{gal} \), which is to be expected since both describe large-scale features across the sky. If we used exactly the same

\(^6\) https://fermi.gsfc.nasa.gov/ssc/data/analysis/scitools/source_models.html
FIG. 4. Posterior distributions for CSiBORG simulation 7444 of the parameters describing the gamma-ray flux in the energy range 30 – 50 GeV. We include templates proportional to the $J$ factor ($A^J_2$) and $D$ factor ($A^D_2$), as well as isotropic ($A^\text{iso}_9$), galactic diffuse ($A^\text{gal}_9$), and point source ($A^\text{psc}_9$) contributions, and define $A_9 \equiv 10^{-22} \text{cm}^{-2} \text{s}^{-1} \text{MeV}^{-1}$. The contours show the 1 and 2$\sigma$ confidence intervals.

FIG. 5. One-dimensional posterior distributions on the coefficients describing the flux proportional to the (upper panel) $J$ factor and (lower panel) $D$ factor in the energy range 30 – 50 GeV. Each black line gives the posterior distribution for an individual CSiBORG simulation, and the red line is the mean of these, i.e., the final posterior distribution.

selection criteria as the Fermi analysis which produced the non-DM templates, then $A^\text{gal}_1$, $A^\text{iso}_1$ and $A^\text{psc}_1$ would all have a mean of unity. This is not true here because the isotropic template is calibrated for latitudes $10^\circ < |\lambda| < 60^\circ$, whereas we fit our template to $|\lambda| > 30^\circ$. In general, we find $A^\text{iso}_i$ to be slightly smaller than 1. This is more prominent in the higher energy bins; we find that $A^\text{iso}_i$ is closer to unity at lower energy. We verify that this is not due to the addition of the $J$ or $D$ factor templates by rerunning the analysis with $A^D_1 = A^J_1 = 0$ and find that $A^\text{iso}_i$ remains less than one.

We generate such MCMC chains for each of the 101 CSiBORG simulations, and we plot the resulting one-dimensional posterior distributions for $A^J_1$ and $A^D_1$ in Fig. 5. Knowing that each CSiBORG simulation is a fair Monte Carlo sample, the final posterior distribution on $A^J_9$ is simply the average of each individual probability distribution, which yields the red lines in the figure. When marginalised over the BORG chain, we again find $A^J_1$ and $A^D_1$ are consistent with zero for this energy bin.

This process is repeated for each energy bin to determine the posterior for a given $A^J_1$ or $A^D_1$, marginalised over all other contributions to the gamma-ray sky and over the uncertainties involved in producing maps of the $J$ and $D$ factors. These spectra are displayed in Fig. 6, where we indicate the maximum posterior points by the circles and 1$\sigma$ confidence intervals by the error bars. For a given DM mass and channel, these posteriors can be trivially transformed into constraints on $\langle \sigma v \rangle$ or $\Gamma$ for a given energy bin. We then simply multiply the posteriors from each bin to determine our final constraint on these parameters.

In Fig. 7, we plot the 95% upper limit on $\langle \sigma v \rangle$ as a function of DM particle mass, $m_\chi$, for an annihilation which solely produces particle-antiparticle pairs of a single type, but for any standard-model quark, charged lepton or gauge boson (except photons). We compare these constraints to the thermal relic cross section ($\langle \sigma v \rangle_{\text{th}} \approx 3 \times 10^{-26} \text{cm}^3 \text{s}^{-1}$), such that, if the curve falls below this value in Fig. 7, then we rule out DM being a thermal relic for the corresponding mass and annihilation channel at 95% confidence. For all annihilations producing gluons or quarks less massive than the bottom quark, we see that, if DM is a thermal relic, it should be more massive than $\sim 7 \text{ GeV}/c^2$ since we rule out smaller masses. We are unable to rule out the thermal relic cross section at any mass for production of bottom quarks, top quarks, W bosons or Z bosons. Our constraints for lepton production are much weaker at a given particle mass, such that our constraints for electron or muon production do not cross $\langle \sigma v \rangle_{\text{th}}$. We rule out $\tau$ production for $m_\chi \lesssim 6 \text{ GeV}/c^2$ at this cross section.
Turning our attention to DM decay, Fig. 8 shows the inferred decay rate, $\Gamma$, for different decay channels as a function of $m_\chi$. Contrary to our analysis of DM annihilation, we find that for the majority of channels we infer a nonzero $\Gamma$ at over $2\sigma$ confidence for at least some $m_\chi$ (corresponding to nonzero $A^D_i$ in Fig. 6). The results are relatively insensitive to the DM particle mass, provided $m_\chi$ is above the threshold for production. For the $b\bar{b}$ channel, we find the inferred $\Gamma$ is $\sim 6 \times 10^{-28} s^{-1}$, which corresponds to approximately one decay per Hubble time in a volume $\sim 280 \text{km}^3$ at mean cosmological density if $m_\chi = 100 \text{GeV}/c^2$. This is around the smallest $\Gamma$ that has been constrained by any previous study (see Section VI C 2). The inferred $\Gamma$ for decay to the lightest charged leptons is approximately an order of magnitude larger than this.

To determine the overall detection significance, we compute the coefficient for the total flux across all energy bins which multiplies the $D$ factor

$$A^D_{\text{tot}} \equiv \sum_i A^D_i \Delta E_i = 1.02^{+0.24}_{-0.28} \times 10^{-16} \text{ cm}^2 \text{s}^{-1}. \tag{53}$$

Simply dividing the best fit value by the lower error would suggest that our detection of a contribution to the gamma-ray sky proportional to the $D$ factor has a significance of $3.6\sigma$ when averaged over all available energies. Since our posterior is non-Gaussian, we wish to compute this significance through other methods. We compute the maximum likelihood for each $A^D_i$ and, since each energy bin is treated as independent, the maximum likelihood for $A^D_{\text{tot}}$ is the product of these values. We compare this to the likelihood for $A^D_i = 0$ and find the change in log likelihood between these two cases is $\Delta \chi^2 = 2\ln \mathcal{L} = 11.2$, which is equivalent to $3.3\sigma$ for a Gaussian likelihood or a change in the Bayesian Information Criterion (BIC) [97] of 9.0, if one takes the $A^D_{\text{tot}} \neq 0$ model as having one more parameter. In Section VIA 2 we ask whether this is due to DM decay, finding that a non-DM spectrum is preferred by the data.

VI. DISCUSSION

In this section, we discuss the possible origin of our results, their limitations and a comparison to existing results in the scientific literature.

A. Interpretation of results

In this section we investigate which objects and observations drive our results and whether there are non-DM explanations for the signal proportional to the $D$ factor.

1. Annihilation

For each channel and $m_\chi$, we compute the change in log likelihood between $\langle \sigma v \rangle = 0$ and the $1\sigma$ constraint on $\langle \sigma v \rangle$ separately for each energy bin to determine which energy range is dominating our constraint. For DM particle masses $m_\chi \lesssim 10 \text{GeV}/c^2$, our constraints are driven by the first three energy bins. For these masses, we expect there to be very few photons produced at high
energies at fixed $\langle \sigma v \rangle$, so these bins are unable to constrain $\langle \sigma v \rangle$ as very large values are required to produce an appreciable flux. As we move to higher masses, we notice the effect of the data at higher energy, such that at the highest masses we find that the sixth energy bin ($6.5 \rightarrow 10.8$ GeV) is the most constraining. We find a similar trend if we compare using the 95% confidence limit instead of the 1σ constraint.

To determine which halos drive our constraints, we produce several $J$-factor maps where each one is only due to objects in a given mass range. We create separate maps for halos in moving bins of width $\Delta \log_{10}(M_{\odot}/M_\odot) = 1$. We rerun the inference for a single representative CSiBORG simulation (9844) separately for each of these mass bins; i.e., we assume that only a single mass bin contributes to the total $J$ factor. We plot the constraint on $\langle \sigma v \rangle$ as a function of halo mass in Fig. 9 and observe that the tightest constraints are obtained for halos in the range $\sim 10^{11.5} \rightarrow 10^{14.5} M_\odot$. If one studied a single object at a fixed distance, then the most massive halo would give the tightest constraints since it has the largest $J$ factor. However, such massive objects are rare, so there is a compromise between having many objects of a given mass across the sky and those objects having a large $J$ factor. Given the tight constraints one can obtain with dwarf galaxies in the Local Group, it is perhaps not surprising that the inclusion of lower mass objects can lead to an improvement in the upper limit on $\langle \sigma v \rangle$. The inclusion of these structures in this work was possible due to the use of constrained simulations, which provide plausible realisations of these halos given the ICs that are constrained on large scales. As a result, Fig. 9 shows the types of objects in the nearby universe one should target to extract maximum information about DM annihilation.

Since we found a nonzero flux proportional to the $D$ factor, we rerun the analysis for all CSiBORG simulations but infer both $A^D_i$ and $A^f_i$ simultaneously. In this way, when making our constraint on $\langle \sigma v \rangle$, we now marginalise over this contribution. Note that in this marginalisation we do not assume a spectral form for the $D$ factor template, such that we marginalise over any source whose spatial distribution is proportional to the local DM density, which may or may not be DM decay. As anticipated in Section V, we find that our constraints become tighter, such that our upper limit on $\langle \sigma v \rangle$ is typically over a factor of 2 smaller. This is as expected: the negative degeneracy between $A^D_i$ and $A^f_i$ (Fig. 4) means that, if we allow $A^D_i > 0$, we must reduce $A^f_i$ so that the total flux from these two contributions is approximately constant. When marginalising over $A^D_i$, this will result in a posterior on $A^f_i$ which is necessarily narrower. Since these results are tighter than when we set $A^D_i = 0$, we choose to report the latter as our fiducial results so that our conclusions are conservative.

\section{Decay}

In Fig. 8 we found a nonzero DM decay rate is compatible with the observed gamma-ray sky at over $2\sigma$ confidence for a range of DM masses and decay channels. More conservatively, one would say that we find a signal which is proportional to the $D$ factor; i.e., the emitted flux from some region appears to be proportional to the local density, and is compatible with the spectrum of DM decay. This source does not necessarily need to be DM decay, which can be investigated by fitting the inferred spectrum to a different model. For this we choose a power-law profile, such that the parameter $A^D_i$ arises from integrating the spectrum

$$\frac{dN}{dE_\gamma} = A_p \left( \frac{E_\gamma}{E_0} \right)^p,$$

across energy bin $i$, where we normalise to $E_0 \equiv 1$ GeV. We place broad, uniform priors on $A_p$ and $p$ in the range $[0, 10^{-18}]$ cm$^{-2}$s$^{-1}$MeV$^{-1}$ and $[-5, 2]$, respectively, and
find
\[
A_p = (4.1 \pm 1.5) \times 10^{-20} \text{ cm}^{-2}\text{s}^{-1}\text{MeV}^{-1},
\]
\[
p = -2.75_{-0.46}^{+0.71},
\]
where the limits are at 1σ confidence. To enable a comparison, we plot the spectrum for this model and a DM decay model in Fig. 6, where we see that the power law fits better at most energies.

To assess the relative goodness of the fit of the two models we compute the BIC. Since we have set deliberately wide priors on our model parameters, ratios of the Bayesian evidence are difficult to interpret. For all channels and masses, we find that the BIC prefers the power-law spectrum, with BIC ≥ 1.3 (the bound is saturated for the gluon channel). Decays via the $tt$ channel are least preferred by the data, with BIC ≥ 4.0. We therefore conclude that, although we do find an excess of gamma-ray flux that traces the density of DM, its spectrum is fit marginally better by a power law, so this is not evidence for DM decay.

This conclusion is consistent with previous works studying the origin of the residual gamma-ray flux. The cross-correlation of the gamma-ray sky with galaxy catalogues has been detected at $2 - 4\sigma$ \cite{98}, and the spectral index of the 1-halo contribution was found to be $-2.7$ if a single power law is assumed, which is consistent with our inferred index and has a similar significance of detection. Moreover, analysis of the angular power spectrum of the gamma-ray sky \cite{99} suggests a component that can be modelled as a double power law with an exponential cutoff, with power-law indices $-2.55 \pm 0.23$ and $-1.86 \pm 0.15$, the former of which is consistent with our result. Ackermann et al. \cite{99} note that this spectrum is compatible with blazarlike sources being the dominant component at these energies. Since we detect a nonzero contribution with the same spatial variation as our $D$ factor maps, our work suggests that these excesses could be due to sources with a linear bias with respect to the local DM density.
B. Systematic uncertainties

In this section we investigate potential systematic errors in our analysis by changing some of the analysis choices in Section IV. For computational convenience, throughout this section we use only one C$\text{S}1\text{BO}R$G realisation (simulation 7444 as given in [58]) unless otherwise stated.

1. Computing the $J$ and $D$ factors

In Section IV.B.1 we computed the $D$ factor and non-halo contribution to the $J$ factor by smoothing simulation particles onto a grid with a kernel inspired by SPH. Besides this kernel, we also consider the CIC density assignment in order to quantify the impact of the kernel choice on our constraints. The median change in the constraint on $\langle \sigma v \rangle$ is 2% for the $bb$ channel if we change to this kernel. For the DM decay inference, we find the median and 95% upper limit on $\Gamma$ change by a median of 16% and 15%, respectively, for this channel. The inferred value of $p$ changes only by 0.01 when we change to the CIC kernel. If the low density regions were driving our constraints, then one would expect large differences between the two procedures, since these regions have the fewest simulation particles and the two kernels have different noise properties for low particle counts. However, we do not see this since the expected flux is highest in the high density regions and the low density regions are relatively unconstraining.

For computational convenience, for our fiducial analysis we chose a HEALPix resolution of $\text{nside}=256$. We rerun the analysis at coarser resolution ($\text{nside}=128$) and find that our constraints on $\langle \sigma v \rangle$ weaken by a median change of 48% across all masses for the $bb$ channel. The median and 95% upper limit on $\Gamma$ change by a median of 13% and 10% respectively for the $bb$ channel, and the inferred value of $p$ changes only by 0.03 when we lower the HEALPix resolution. It is unsurprising that the $D$ factor analysis is less affected by this choice; for the $J$ factor our template is dominated by high density peaks in the DM density field, since the flux is proportional to the square of the density. By using a higher resolution map, one can localise these peaks better to obtain tighter constraints if these are not aligned with peaks in the observed gamma-ray sky.

The $J$ and $D$ factor maps were calculated for each of the 101 C$\text{S}1\text{BO}R$G simulations. By utilising the full suite, we marginalise over the uncertainties in the constrained density modes from both the B$\text{O}R$G algorithm and the unconstrained, small-scale modes. To verify that we have a sufficiently large number of simulations to achieve this, we rerun our analysis 100 times for the $bb$ channel using 50 randomly selected simulations to determine a bootstrap uncertainty on our constraints. The standard deviation of the 95% upper limit on $\langle \sigma v \rangle$ has a median value of 19% when considering all masses. The inferred $\Gamma$ has a median bootstrap uncertainty of 4%. The uncertainty on the inferred power-law index, $p$, is 0.05 and the fractional bootstrap uncertainty on $A_p$ is 4%, which are small compared to the uncertainties we quote in Eq. 55. We therefore conclude that the number of constrained simulations is adequate.

2. Halo density profile

After identifying halos within the C$\text{S}1\text{BO}R$G simulations, we assumed that all halos have NFW profiles with masses as given by the halo finder and concentrations given by the mass-concentration relation of [88]. To determine the sensitivity of our constraints to the assumed profile, we rerun the analysis but assuming that all halos are described by Einasto profiles, as calculated in Section IV.B.2. We find our constraints on $\langle \sigma v \rangle$ can be up to 80% tighter if one uses an Einasto profile compared to a NFW. For small mass halos we find that the Einasto profile leads to larger densities near the centre than the NFW profile and, given the importance of these lower-mass objects for our constraint (Section VI.A.1), this leads to smaller values of $\langle \sigma v \rangle$. Note that a similar tightening of the constraints was observed in [100] when the Milky Way profile is changed from a NFW to an Einasto profile. We choose to report the most conservative constraints, hence the choice of NFW profiles in our fiducial analysis.

Both the form and parameters of the NFW and Einasto profiles are inspired by N-body, DM-only simulations. Observations and hydrodynamical simulations suggest that these may not accurately describe the true density profiles. In the presence of baryons, the DM profile could be steeper due to adiabatic contraction during galaxy formation [101, 102], or shallower due to the subsequent stellar feedback (e.g. [103, 104]). For steeper slopes of the density profile, the $J$ factor near the centre would also be larger, and hence one would expect tighter constraints on $\langle \sigma v \rangle$. Ideally one would use profiles motivated by hydrodynamical simulations; however, common parameterisations such as [105] apply primarily to sub-Milky Way mass halos while most of those produced in C$\text{S}1\text{BO}R$G are in the group and cluster regimes. Rather than perform a large extrapolation, we leave it to future work to implement a robust baryonification scheme on these scales.

3. Substructure uncertainties

Since the gamma-ray flux from DM annihilation is proportional to the square of the density, the substructure of DM halos is an important contribution that one must consider; if one computes the angular power spectrum for the $J$ factor, $C_\ell$, one finds that $(2\ell+1)C_\ell$ is approximately constant at the smallest scales considered in this work. Usually this substructure is modelled as a mass-dependent multiplicative boost factor [e.g. 106] and
uncertainties captured by looking at the extreme values of the boost for given masses [51]. We, on the other hand, capture substructure and its uncertainty through CLUMPY’s probabilistic approach to substructure modelling. This led us to use a non-Poisson likelihood, since we introduced uncertainties, $\sigma_{jp}$, on the Poisson means. To evaluate the impact of this choice, we rerun the analysis assuming a Poisson likelihood by setting $\sigma_{jp} = 0$. We find that our constraints typically change by a few percent across all channels and masses, indicating that the impact of this uncertainty is negligible. However, we note that this could not have been known a priori. Although the fractional uncertainties are small near the centres of halos, this is not true in the outskirts, motivating our thorough treatment of uncertainties.

For simplicity, we previously neglected the uncertainty that arises due to stochasticity in the mass-concentration relation. We find our constraints are not very sensitive to the scatter in this relation. For the $bb$ channel across all masses, the median change in the constraint on $\langle \sigma v \rangle$ is 1.4% if this uncertainty is included.

Another source of systematics due to substructure modelling might be driven by the breakdown of our assumption that the uncertainty in the $J$ factor is Gaussian. Namely, as the considered mass of the clump grows, the total number of such clumps within the host halo decreases. Therefore, it is expected that at some point we transition from the Gaussian into a Poisson regime [107]. Furthermore, it is not obvious that the contribution to the $J$ factor from these more massive clumps will not outshine the cumulative contribution of the lower-mass clumps. To check for this, we use the -h5 module of the CLUMPY package to explicitly draw substructure realisations for a typical halo ($M_h \approx 5 \times 10^{13} M_\odot$) from the CSiBORG simulations. We modify Eq. 35 such that

$$\langle J_p \rangle = \langle J_{\text{cont},p}(M_{th}) \rangle + \langle J_{\text{drawn},p}(M_{th}) \rangle,$$

and

$$\langle J_{\text{cont},p}(M_{th}) \rangle = J_{\text{sm},p} + \langle J_{\text{subs},p}(M_{th}) \rangle + \langle J_{\text{cross},p}(M_{th}) \rangle,$$

where we introduce

$$\langle J_{\text{drawn},p}(M_{th}) \rangle = \frac{1}{N_{ds}} \sum_{i} J_{\text{drawn},pi}(M_{th}),$$

with $N_{ds}$ being the total number of explicit realisations of the clumps with a mass above a given mass threshold, $M_{th}$. The quantities $\langle J_{\text{subs},p}(M_{th}) \rangle$ and $\langle J_{\text{cross},p}(M_{th}) \rangle$ from Eq. 57 are obtained by replacing the upper limit of the clump mass distribution by $M_{th}$, i.e., replacing $M_2$ with $M_{th}$ in Eq. 27. To estimate $\langle J_{\text{drawn},p}(M_{th}) \rangle$, we run $N_{ds} \approx 1000$ explicit realisations of substructure clumps for a typical halo, requiring that we capture fluctuations in the value of $J_{\text{sm},p}$ – the leading contribution to the total $J$ factor of the halo – at the percent level. In other words, any clump whose contribution to the given pixel will induce a fluctuation to the value of $J_{\text{sm},p}$ of the order of $\sim 1\%$ will be explicitly drawn onto the HEALPix grid. This is equivalent to taking one sample from Eq. 21, but with a modified mass range of the mass function, and setting the lower limit for this draw to be $M_{th}$. For more details see section 2.4.3 of [76].

For this experiment, we selected a HEALPix resolution of $nside = 1024$, corresponding to the Fermi-LAT angular resolution. The corresponding threshold mass for this setup translates to $M_{th} = 5.3 \times 10^9 M_\odot$ for our chosen halo. We find that

$$\langle J_{\text{drawn},p}(M_{th}) \rangle \sim 0.04 \langle J_{\text{cont},p}(M_{th}) \rangle,$$

which justifies our starting assumption of treating the substructure contribution to the total $J$ factor as Gaussian, since the “drawn” (Poisson) component is subdominant compared to the “continuous” (Gaussian) contribution.

Note, however, that choosing a smaller $M_{th}$, i.e., looking at even smaller fluctuations of $J_{\text{sm},p}$, would lead to probing the even lower-mass end of the substructure mass function, which would, of course, alter the ratio of the “drawn” and “continuous” components. However, going below this limit would already enter into a regime where drawing $10^4 - 10^5$ clumps from the corresponding version of Eq. 21 will be necessary, which is computationally expensive and well within the validity of the Gaussian approximation. As a comparison, using $M_{th} \approx 5.3 \times 10^9 M_\odot$ required around $\sim 10^4$ draws. Throughout this section we assumed the same DM profile parameterisation and mass-concentration relation as in the fiducial inference (see Section VIB). The conclusions are unchanged for the same halo using the Einasto profile.

4. Non-DM templates

The point source template is designed to remove small-scale emission so that large-scale variations in the gamma-ray sky can be modelled more robustly. Since this template is derived from the data, there is a risk that a subset of the point sources could be due to annihilating or decaying regions of high DM density. Modelling these as a non-DM component would therefore be incorrect. To assess this, we compute the angular cross-correlation function between halos identified from the CSiBORG simulations and the positions of point sources detected by Fermi-LAT. We find no significant correlation at any scale, justifying our modelling assumptions. This would be expected from the lack of degeneracy between the amplitude of the point-source template, $A_{psc}^i$, and the amplitudes of the $J$ and $D$ factor templates, $A_j^i$ and $A_D^i$, in Fig. 4. This suggests that our constraints are not driven by point sources, so the precise model we use for these is not important.

Although in Fig. 4 we see there is little degeneracy between the parameters describing galactic diffuse emission and the DM annihilation or decay parameters, one should
verify that the results are robust to reasonable variations in these non-DM templates. In our fiducial analysis we used the most recent galactic diffuse model provided by the Fermi Collaboration (gll_iem_v07). We rerun the analysis with an older model (gll_iem_v02) and find our constraints on \( \langle \sigma v \rangle \) are slightly weaker, with a median change of 9% across all mass bins for the bb channel, and that the inferred value of \( \Gamma \) can vary by \( \sim 35\% \). Although there is some variation as we change the model, these are at a similar level to other systematic effects. The inferred power-law index, \( p \), only changes by 0.01 so it is insensitive to this choice.

In our analysis we assume that the amplitudes of the non-DM templates are independent, and thus we do not impose a prior on the shape of the spectrum, i.e. on the relative amplitudes between energy bins. If one were to assume that the shape of the spectrum was identical to that in the Fermi analysis, then one would enforce the same amplitude for a given template across all bins. If these parameters were strongly degenerate with \( \sigma \) and/or \( A_D \), or if their values varied significantly with energy, then imposing such a prior could lead to tighter constraints on \( \langle \sigma v \rangle \) and \( \Gamma \). Neither of these criteria apply (see Fig. 4), we do not make such a choice, so the spectra of the astrophysical templates are determined empirically. Similarly, we have not imposed the spectra of gamma-rays expected from annihilation or decay, instead deriving constraints from each energy bin separately.

5. High redshift sources and optical depth

The gamma rays emitted from either DM annihilation or decay would interact with the extragalactic background light (EBL) or CMB photons [108, 109]. This interaction manifests itself through pair production and/or annihilation or decay. For the redshift range considered in this paper (\( z \lesssim 0.05 \)), the attenuation of the photon flux due to interaction with background photons will not be significant, except at very high energies (\( \sim \) TeV), which lie well above the maximum photon energies we consider here (\( \sim 50 \) GeV), and thus we neglected this contribution.

Although this is the case for the very nearby Universe, there is also a contribution to \( J \) and \( D \) from sources outside the CSiBORG volume. The expected contribution to the differential photon flux from annihilation is [7]

\[
\left\langle \frac{d^2\Phi_{\gamma}^{\text{ann}}}{dE_\gamma d\Omega} \right\rangle = \frac{\langle \sigma v \rangle \rho_{DM,0}^2}{8\pi m_\chi^2} \int dz \left( \frac{dN_\gamma}{dE_\gamma} \right)_{E_\gamma = E_\gamma(z(1+z))} \times \frac{(1+z)^3}{H(z)} e^{-\tau(E_\gamma,z)} \langle (1 + \delta(z, \Omega))^2 \rangle, \tag{60}
\]

where \( \rho_{DM,0} \) represents present-day DM density and \( \delta(z, \Omega) \) is the density fluctuation. This can be directly computed from the nonlinear matter power-spectrum (see, for example, [110]) or by using the halo model approach [49, 111]. Within the halo model, this factor comes directly from averaging the one-halo annihilation luminosity over the halo mass function. This is equivalent to marginalising over plausible realisations of the DM distribution in our Universe by utilising the Press-Schechter [112] formalism, or any other halo-formation model.

As in [111] (see Fig. 10 in their Appendix B), we estimate the integrand at a given \( E_\gamma \) and integrate between \( z = 0.05 \) and \( z = 10 \) to determine the ratio of this contribution to that explicitly modelled from CSiBORG. We find that this ratio is approximately unity at \( E_\gamma = 5 \) GeV for \( m_\chi = 10 \) GeV/c\(^2\) for the bb channel. One may be concerned that this is an important contribution; however, since our chosen HEALPix resolution of \( \text{nside} = 256 \) corresponds to a physical scale of \( \sim 0.6 h^{-1} \) Mpc at the edge of the CSiBORG volume, one would expect that the extragalactic sources beyond \( z \gtrsim 0.05 \) are unresolved, and therefore this contribution will almost entirely be absorbed into the isotropic template. Of course, clustering of sources at redshifts beyond the CSiBORG volume would lead to an anisotropy in this unresolved emission. Our constraints are completely independent of how one models the isotropic part of the high-redshift component, and we leave it to further work to model the fluctuations about this, for example, by including constrained realisations of the density field for larger volumes. Since this contribution can only increase the \( J \) factor, we always underestimate the \( J \) factor in our templates, making our constraints conservative.

We note that in Eq. 60 one must correct for the redshift of emission: i.e., the spectrum should be evaluated at \( E_\gamma(z(1+z)) \) for a source at redshift \( z \) if we observe a photon at energy \( E_\gamma \). Since we only considered sources at \( z \lesssim 0.05 \), we neglected this effect. If we consider the extreme case where all our sources were actually at \( z = 0.05 \), we would find that our estimate of the flux at a given \( \langle \sigma v \rangle \) is correct to within \( 5\% \) in the five lowest energy bins for the bb channel at \( m_\chi = 100 \) GeV/c\(^2\) (similar effects are seen for other channels and masses). Since this is comparable to the size of other reasonable variations to the model and this is an unrealistically extreme case, we are justified in making this assumption.

C. Comparison to literature

To enable a comparison between the constraints on \( \langle \sigma v \rangle \) and \( \Gamma \) obtained in this work using large-scale structure and those from the literature, we now briefly summarise other methods for inferring these parameters and the results they produce.
1. Annihilation

The release of energy by annihilating DM has the potential to affect several observables over the Universe’s history. The earliest important observable is the element yield from Big Bang Nucleosynthesis (BBN): a 100 GeV/c^2 DM particle with the thermal relic cross section would release 1 MeV of energy for every baryon in the Universe per Hubble time during the BBN era. This has the potential to alter subdominant nuclear reactions, although the effect is not strong enough to lead to stringent constraints [113, 114]. The next important epoch is recombination, where annihilating DM has the potential to ionise a non-negligible fraction of the hydrogen in the Universe. This would absorb CMB photons after recombination, to which the CMB angular power spectra and power spectra are acutely sensitive [115]. This allows thermal relic DM with a velocity-independent cross section to be ruled out for masses below \( \sim 10 - 30 \text{ GeV/c}^2 \) depending on the annihilation channel [46, 116]. Of course, these bounds could be evaded by a large branching fraction into neutrinos or other particles with no electromagnetic interaction. The CMB constraints are particularly important for light DM \( (m_\chi \lesssim 1 \text{ GeV/c}^2) \) where the effective area and angular resolution of telescopes such as Fermi-LAT are poor. At even later times \( (2 \lesssim z \lesssim 6) \), observations of the Lyman-\( \alpha \) forest constrain the gas temperature (e.g., [117, 118]), which would be increased by annihilations, although this has not been used to set quantitative constraints.

These bounds were derived purely by considering the effects of energy injection into the Universe, but more information is available from observations of the potential annihilation products themselves. This is most often done by means of high-energy photons (a common final product regardless of annihilation channel), and forms the context for our own analysis. Of course, the spectra of the final-state photons depend crucially on the channel, as we have described previously. In the local Universe, the most promising targets are the galactic centre, nearby groups or clusters, and dwarf galaxies in the Local Group. The former is the greatest nearby concentration of DM, but also suffers from large astrophysical backgrounds, and the expected signal depends sensitively on the poorly known DM density profile of the Milky Way. Nevertheless, there are claims for a gamma-ray excess that could be due to annihilating DM [12]. In particular, Hooper and Goodenough [12] claim the excess is well fit for \( m_\chi \sim 7 - 10 \text{ GeV/c}^2 \) with \( \langle \sigma v \rangle \sim (0.5 - 5) \times 10^{-26} \text{ cm}^3\text{s}^{-1} \) annihilating via the \( \tau^+ \tau^- \) channel. Although we cannot rule out the lower values of \( \langle \sigma v \rangle \) for this mass range, we do find that \( \langle \sigma v \rangle < 4.5 \times 10^{-26} \text{ cm}^3\text{s}^{-1} \) at 95\% confidence for these masses and this channel, which is incompatible with the larger values of \( \langle \sigma v \rangle \) reported. Clusters are also massive accumulations of DM and permit a statistical analysis, but also suffer from potentially significant backgrounds. Dwarf galaxies, although smaller and less dense, have a much lower baryonic mass and hence the lowest contribution from degenerate astrophysical effects, affording a cleaner test. However, one is limited to a small sample size and thus one has to assume that the objects do not have peculiarities, e.g. unusual boost factors. By looking at a larger number of sources, as we do here, one can average over a more representative sample of substructure.

Fermi-LAT has been used to set limits on the annihilation cross section using dwarf galaxies, the Milky Way halo [54, 119, 120], and galaxy groups [45]. The strongest constraints come from the dwarfs, which, due to their lower distances, offer higher peak signals than clusters [121]. These have been used to rule out the thermal relic cross section for masses below \( \sim 100 \text{ GeV/c}^2 \) assuming annihilation to \( b \) quarks [42, 43], as depicted in Fig. 7 (although see [122]). Even stronger constraints, ruling out the thermal relic scenario to \( O(\text{TeV}) \) mass scales for annihilation to \( bb \) have been claimed from a radio search of the Large Magellanic Cloud [123]. Somewhat weaker constraints have also been obtained using dwarf irregular galaxies [44] and by cross-correlating Fermi-LAT data with the positions of nearby galaxies without knowledge of those galaxies’ distances (dashed line in Fig. 7) [39, 40]. Further information can be gleaned by cross-correlating gamma-ray flux with a tracer of density such as gravitational lensing [41].

Data from the ground-based air Cherenkov telescopes VERITAS, MAGIC, HAWC, and H.E.S.S. have also been used to set constraints from dwarfs, which dominate those from Fermi-LAT for \( m_\chi \gg 1 \text{ TeV/c}^2 \) [124–127]. H.E.S.S. has also been applied to the galactic centre, achieving stronger constraints at very high energies at the cost of increased systematic uncertainty due to astrophysical backgrounds [100, 128]. It is worth noting also that constraints on both annihilation and decay can be set by direct detection laboratory experiments, although these are considerably weaker than astrophysical constraints [129].

Annihilating DM produces other cosmic rays besides photons, most notably positrons and antiprotons. The AMS-02 instrument has provided data on the spectrum of a wide range of cosmic ray species [130–132]. Despite uncertainties due to cosmic ray propagation and the impact of the Sun’s magnetic field, antiproton observations have been used to set bounds that beat those from Fermi-LAT in some cases, for example, constraining the \( \mu^+ \mu^- \) channel to \( m_\chi \sim 100 \text{ GeV/c}^2 \) at the thermal relic cross section [131]. It is also possible to search for gamma-ray lines, which are generically expected to be weak but may be prominent if the DM particle decays to charged particles similar to it in mass. Line limits from Fermi-LAT and H.E.S.S. are presented in [100, 133] and [134, 135] respectively.
Similar considerations to those of Sec. VI C 1 allow cosmological constraints to be placed on DM decay. These constraints are stronger at lower redshifts, where a greater fraction of DM decays per Hubble time for fixed decay rate. This allows BBN to test decay lifetimes around $10^{13}$ s, the CMB $10^{25}$ s, and the Lyman-$\alpha$ forest $10^{25} - 10^{26}$ s [7].

Constraints can also be derived from gamma-ray and neutrino telescopes. In particular, data from Fermi-LAT, AMS-02, PAO, KASCADE, and CASAMIA have been used to constrain the DM lifetime at the $10^{27} - 10^{28}$ s level for $10^2 \text{GeV} / c^2 < m_x < 10^{17} \text{GeV} / c^2$ [136, 137]. For lower-mass DM decaying primarily leptonically, bounds at the $10^{25} - 10^{26}$ s level can be set from X-ray and gamma-ray telescopes, the spectrometer on board the Voyager I spacecraft, and the heating of gas-rich dwarf galaxies, as well as the Lyman-$\alpha$ forest and CMB as described above [47, 138]. These constraints imply that over a very large DM mass range, only a tiny fraction of the total DM can decay during the lifetime of the Universe. Decaying DM can also be constrained using the masses and abundances of Milky Way satellites in case the decay gives momentum to the DM particle, which provides a constraint of order the age of the Universe (e.g., [139]). The inferred values of $\Gamma$ in this work are compatible with these constraints.

D. Future directions

1. Including additional mass

Our analysis deliberately targets large-scale structure as a source of annihilation or decay flux in order to be fully complementary to studies of particular objects while avoiding their systematics. This has made our constraints conservative because significant $J$ and $D$ factor contributions come from the Milky Way halo and dwarf spheroidals in the Local Group. Incorporating these into our mass model would therefore produce the most constraining results possible, modelling flux from all mass in the local Universe. For the Milky Way this could be done by detailed modelling of the properties of our host halo along the lines of [119, 120] but using the latest data from Gaia [140, 141]; this will be the subject of future work. A separate likelihood component could be added for local dwarf galaxies (cf. [42, 43]). We note that inferring ICs which could produce such structures with the correct masses and locations using a process similar to the SIBELIUS simulations [142, 143] would be a computationally demanding task. A more feasible approach may be to populate larger halos with such objects a posteriori in a manner similar to how one paints galaxies onto a N-body simulation.

2. Velocity dependence

In this work we have assumed that $\sigma v$ is independent of energy. One can generalise this such that the cross section is multiplied by a function $S$ of the relative velocity between two DM particles, $v_r$, i.e. $(\sigma v) = (\sigma v)_0 S(v_r / c)$. The velocity-dependent term is commonly modelled as $S(x) = x^n$, where in this work we have considered $n = 0$ ($s$-wave) scattering. Other velocity dependencies are theoretically interesting: in models with minimal flavour violation, $n = 2$ ($p$-wave) annihilation dominates for Majorana fermions forming Standard Model fermion-antifermion pairs, since the $s$-wave is chirality suppressed [144]. A null result for $p$-wave annihilation in the galactic centre is presented in [145]. Similarly, $n = 4$ ($d$-wave) dominates in such models if DM is instead a real scalar [146, 147]. Because of the small DM velocities within halos, one would expect these signals to be harder to detect than $s$-wave scattering. However, the larger velocity dispersion in massive objects increases the relative importance of high mass objects relative to local ones [148], making extragalactic halos interesting targets to probe velocity-dependent DM annihilation. We will analyse these models using our constrained simulations in a future publication.

Furthermore, if DM has long-range self-interactions, then the annihilation is Sommerfeld enhanced [149], corresponding to $n = -1$ and thus a high, potentially detectable annihilation rate. When studying these velocity-dependent cross sections, one cannot use the $J$ factor given in Eq. 5, but instead [150]

$$J_S = \int ds \, d^3 v_1 \, d^3 v_2 \, S\left(\frac{|v_1 - v_2|}{c}\right) f(r, v_1) f(r, v_2),$$

(61)

where $f(r, v)$ is the distribution function of DM particles. This thus requires one to know or model the velocities of DM particles within halos and is therefore left to future work, although we note that $J_S$ has recently been calculated for a range of DM density profiles [151].

VII. CONCLUSIONS

Indirect detection of DM annihilation or decay through gamma-ray emission has previously typically involved inference from a small number of nearby, DM rich objects (the Milky Way, dwarf spheroidals in the Local Group or local groups and clusters) or by cross-correlating the gamma-ray background with other catalogues. Instead, in this work we utilise the CSIBORG suite of constrained simulations of the local $155 h^{-1} \text{Mpc}$ to forward model the predicted gamma-ray sky for $s$-wave DM annihilation or decay due to the large-scale structure. We marginalise over uncertainties in the density field reconstruction, unresolved substructure, and non-DM contributions to the signal, and compare to data from Fermi-LAT via a
MCMC algorithm. We rule out the thermal relic cross section at 95% confidence for DM particles of mass $m_\chi \lesssim 7 \text{ GeV}/c^2$ whose annihilation produces gluons or quarks less massive than the bottom quark. Our constraints for the production of charged leptons are approximately an order of magnitude less stringent, and we are unable to rule out the thermal relic cross section for the production of top or bottom quarks in our fiducial analysis. We infer at $3.3\sigma$ a nonzero contribution to the gamma-ray sky with the same spatial distribution as predicted by DM decay. For the decay to quarks, this corresponds to a decay rate of $\Gamma \approx 6 \times 10^{-28} \text{s}^{-1}$. However, we find that a power-law spectrum is preferred by the data, and we infer that the power-law index is $p = -2.75^{+0.71}_{-0.46}$. If we marginalise over the contribution with the same spatial distribution as DM decay, we obtain constraints on $\langle \sigma v \rangle$ which are twice as tight as our fiducial analysis.

Our constraints on the annihilation cross section are less stringent than those obtained by studying other objects, such as the GCE or dwarf galaxies in the Local Group. Given the sensitivity of the dwarf spheroidal analysis to the prior on galaxy mass [122] and the conflicting explanations for the GCE, this work provides a useful independent probe of novel DM properties by forward modelling the whole gamma ray sky and will thus be sensitive to different systematics. The field-based framework we develop implicitly incorporates not just the two point correlation function – a more traditional way to constrain DM properties from large-scale structure – but all other higher order statistics as well. Since both DM annihilation and decay fluxes are determined by line-of-sight integrals of the density field, the use of constrained simulations provides a convenient way of calculating these integrals for the observed Universe. Future work should be dedicated to a joint inference where one combines the contribution to the DM annihilation or decay signal from large-scale structure with objects from the Local Group. As with analyses on smaller scales, there is some sensitivity to how one parameterises the halo density profiles, and thus future analysis should include procedures to (probabilistically) model baryonic effects on the DM density profile.

**DATA AVAILABILITY**

This work required modifications to the CLUMPY package in order to model substructure in extragalactic halos and process all the halos from CSiBORG at once. These modifications will be included in the next CLUMPY release. Our $J$ and $D$ factor maps may be useful for other indirect detection analyses of current or future data, so we make these maps publicly available at https://cloud.aquila-consortium.org/s/csiborg_jdmaps.

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Some of the results in this paper have been derived using the HEALPy and HEALPix [67, 68], CLUMPY [76–78], EMCEE [93], and FERMIpy [66] packages.

Supporting research data are available on reasonable request from the corresponding authors.

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Appendix A: Numerical Implementation of Likelihood

When evaluating the likelihood in Eq. 44, we often care about small fractional errors on the model, such that the parameter

\[ x \equiv \frac{(\mu - \sigma^2)^2}{2\sigma^2} \]  

(A1)

is large, where we have omitted indices for clarity. We therefore must evaluate \(1F_1(a, c, x)\) for \(x \gg 1\), although this is problematic since \(1F_1\) has the asymptotic expansion [152]

\[ 1F_1(a, c, x) \sim \frac{\Gamma(c)}{\Gamma(a)} e^{x} x^{a-c}, \]  

(A2)

and thus diverges as \(x \to \infty\). We therefore define the function

\[ f(a, c, x) \equiv e^{-x} x^{c-1} 1F_1(a, c, x), \]  

(A3)

which is finite as \(x \to \infty\). Thus the likelihood for a single pixel becomes

\[
\mathcal{L}(\mu, \sigma^2) = \frac{(\mu - \sigma^2)^n}{\sqrt{\pi n} n!} \exp\left(\frac{\sigma^2}{2} - \mu\right) \left(1 + \text{erf}\left(\frac{\mu}{\sqrt{2\sigma^2}}\right)\right)^{-1} 
\times \left(\Gamma\left(\frac{1+n}{2}\right) g_n(x) + 2\Gamma\left(1+\frac{n}{2}\right) h_n(x)\right),
\]  

(A4)

where we have defined the functions

\[ g_n(x) \equiv f\left(\frac{1+n}{2}, \frac{1}{2}, x\right), \]  

(A5)

and

\[ h_n(x) \equiv f\left(1+\frac{n}{2}, \frac{3}{2}, x\right). \]  

(A6)

Given the asymptotic expansion of \(1F_1\), it is clear that Eq. A4 is equivalent to the Poisson distribution in the limit \(\sigma \to 0\) at fixed \(\mu\).

If \(n\) is even, then \(g_n(x)\) can be expressed as

\[ g_n(x) = \sum_{m=0}^{n/2} \alpha_m^{(n)} x^{-m}, \]  

(A7)

where the coefficients obey the recurrence relation

\[ \alpha_m^{(n+2)} = \frac{2}{n+1} \left( n - m + \frac{3}{2} \right) \alpha_m^{(n)} + \alpha_{m-1}^{(n)} \]  

(A8)

and

\[ \alpha_0^{(0)} = 1. \]  

(A9)
Similarly, if \( n \) is odd, then
\[
g_n(x) = \sum_{m=0}^{(n-1)/2} \beta_m^{(n)} x^{-m}, \tag{A10}
\]
for
\[
\beta_m^{(n+2)} = \frac{2}{n+2} \left( n - m + \frac{3}{2} \right) \beta_m^{(n)} + \beta_m^{(n)} , \tag{A11}
\]
and
\[
\beta_0^{(1)} = 1. \tag{A12}
\]
For odd \( n \), the function \( g_n(x) \) has a slightly more complicated expression
\[
g_n(x) = \frac{e^{-x}}{\sqrt{x}} \sum_{m=0}^{(n-1)/2} \gamma_m^{(n)} x^{-m} + \sqrt{\pi} \text{erf} \left( \frac{\sqrt{x}}{2} \right) \sum_{m=0}^{(n-1)/2} \delta_m^{(n)} x^{-m}, \tag{A13}
\]
where
\[
\gamma_m^{(n+2)} = \frac{2}{n+1} \left( \delta_m^{(n)} + (n - m + 1) \gamma_m^{(n)} \right), \tag{A14}
\]
\[
\delta_m^{(n+2)} = \frac{2}{n+1} \left( \delta_m^{(n)} + (n - m + \frac{3}{2}) \delta_m^{(n)} \right), \tag{A15}
\]
\[
\gamma_0^{(1)} = \delta_0^{(1)} = 1. \tag{A16}
\]
Finally, the function \( h_n(x) \) can be expressed in a similar form for even \( n \),
\[
h_n(x) = \frac{e^{-x}}{\sqrt{x}} \sum_{m=0}^{n/2} \epsilon_m^{(n)} x^{-m} + \sqrt{\pi} \text{erf} \left( \frac{\sqrt{x}}{2} \right) \sum_{m=0}^{n/2} \zeta_m^{(n)} x^{-m}, \tag{A17}
\]
where
\[
\epsilon_m^{(n+2)} = \frac{2}{n+2} \left( \zeta_m^{(n)} + (n - m + 1) \epsilon_m^{(n)} \right), \tag{A18}
\]
\[
\zeta_m^{(n+2)} = \frac{2}{n+2} \left( \zeta_m^{(n)} + (n - m + \frac{3}{2}) \zeta_m^{(n)} \right), \tag{A19}
\]
\[
\epsilon_0^{(0)} = 0, \tag{A20}
\]
\[
\zeta_0^{(0)} = \frac{1}{2}. \tag{A21}
\]
Whenever \( x > 1 \), we evaluate the likelihood using Eq. A4 and the series expansions for \( g_n(x) \) and \( h_n(x) \) given above. Otherwise, the likelihood is evaluated using Eq. 44 directly, using the usual series expansion for \( \text{}_{1} F_{1} \)
\[
\text{}_{1} F_{1} (a, c, x) = \sum_{m=0}^{\infty} \frac{(a)_m x^m}{(c)_m m!}, \tag{A22}
\]