Clinical Value of Consensus Molecular Subtypes in Colorectal Cancer: A Systematic Review and Meta-analysis

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Abstract

Background: The consensus molecular subtypes (CMSs) of colorectal cancer (CRC) capture tumor heterogeneity at the gene-expression level. Currently, a restricted number of molecular features are used to guide treatment for CRC. We summarize the evidence on the clinical value of the CMSs.

Methods: We systematically identified studies in Medline and Embase that evaluated the prognostic and predictive value of CMSs in CRC patients. A random-effect meta-analysis was performed on prognostic data. Predictive data were summarized.

Results: In local disease, CMS4 tumors were associated with worse overall survival (OS) compared to CMS1 (hazard ratio [HR] = 3.28, 95% confidence interval = 1.27 to 8.47) and CMS2 cancers (HR = 2.60, 95% confidence interval= 1.93 to 3.50). In metastatic disease, CMS1 consistently had worse survival than CMS2-4 (OS HR range = 0.33 to 0.55; progression-free survival HR range = 0.53 to 0.89). Adjuvant chemotherapy in stage II and III CRC was most beneficial for OS in CMS2 and CMS3 (HR range = 0.16 to 0.45) and not effective in CMS4 tumors. In metastatic CMS4 cancers, an irinotecan-based regimen improved outcome as compared to oxaliplatin (HR range = 0.31 to 0.72). The addition of bevacizumab seemed beneficial in CMS1, and anti-EGFR therapy improved outcome for KRAS wildtype CMS2 patients.

Conclusions: The CMS classification holds clear potential for clinical use in predicting both prognosis and response to systemic therapy, which seems to be independent of the classifier used. Prospective studies are warranted to support implementation of the CMS taxonomy in clinical practice.

KEY WORDS
Colorectal cancer; tumor biology; consensus molecular subtypes; prognostic value; predictive value; treatment; chemotherapy; targeted therapies; personalized medicine
Colorectal cancer (CRC) is the third most common cancer and the second-leading cause of cancer deaths worldwide (1). CRC is a highly heterogeneous disease with respect to clinical and biological features, resulting in striking differences in disease progression and treatment response (2). This diversity of CRC complicates estimation of prognosis and the optimal timing and selection of treatment regimens for individual patients. Currently, mainly pathological staging by TNM, sidedness and a few molecular markers, including mismatch repair (MMR) status, RAS and BRAF mutation status, are regularly used in clinical practice to select patients for specific therapies. Most recently, the BEACON regimen (encorafenib (BRAF-inhibitor), binimetinib (MEK-inhibitor), and cetuximab (anti-EGFR)) was found to be effective for BRAF mutant metastatic CRC (mCRC), and immunotherapy (pembrolizumab; anti-PD-1) significantly prolonged progression-free survival in the first-line setting of patients with microsatellite instable (MSI) tumors (3,4). However, a clinically significant number of patients within the currently used subgroups for treatment selection do not benefit from the available regimens. For instance, 59% of RAS/BRAF wildtype (wt) patients with left-sided tumors, for whom anti-EGFR treatment is an option, do not respond (stable or progressive disease) (5). Therefore, additional molecular tumor characteristics are needed to further personalize therapies and prevent both under- and overtreatment of CRC patients.

The consensus molecular subtype (CMS) classification is a thoroughly studied and robust stratification strategy for CRC (6). The CMS taxonomy was principally founded on the basis of differences in tumor biology rather than clinical outcomes, thereby capturing the intrinsic biomolecular heterogeneity of CRC. Based on differential gene expression in tumor tissue, comprising both cancer cells as well as the microenvironment, CRC can be divided into four subtypes (CMS1-4). CMS1 is the
immunogenic subtype, enriched for MSI tumors and \textit{BRAF}-mutations; CMS2 has epithelial characteristics with marked WNT and MYC signaling and high CIN; CMS3 also has epithelial features but less CIN, is enriched for \textit{KRAS} mutations, and presents with evident metabolic dysregulation; CMS4 is the mesenchymal subtype with prominent TGF-\beta activation, stromal invasion, angiogenesis and an inflammatory, immunosuppressive phenotype (6,7).

Although the CMS taxonomy provides valuable insight into tumor biology and could be used to guide drug development, the direct clinical utility of the CMS classification lies in the possibility to estimate survival (prognostic value) and select patients for both adjuvant or palliative chemotherapy and currently used targeted agents (predictive value). The poor prognostic value of CMS4 and the relatively favorable prognosis of CMS1 and CMS2 in non-metastatic disease have been demonstrated before (6), although no aggregated evidence of different survival measures per CMS was previously reported. We argue that molecular stratification of CRC patients in clinical trials should be encouraged and that the path towards clinical implementation should be investigated. We therefore conducted a systematic review and meta-analysis to evaluate the current state of the evidence on the prognostic and predictive potential of the CMSs to guide future treatment strategies. Both in local and metastatic disease, we compared the CMSs for overall survival (OS), progression-free survival (PFS), relapse-free survival (RFS) and survival after relapse (SAR) and we assessed benefit of (adjuvant) chemotherapy regimens and targeted agents in each subtype.

\textbf{Methods}
Study protocol and objectives

The PROSPERO international prospective register of systematic reviews was consulted before the start of the study and no ongoing reviews with similar scope were identified. Our study protocol was submitted to PROSPERO before the start of data extraction and is accessible via the PROSPERO database (PROSPERO identifier CRD42020165483) (8). Reporting is in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines (9).

The primary objectives of this study were to evaluate the prognostic and predictive impact of the CMS classification in CRC patients, hereby investigating their relevance for use in clinical practice.

Search strategy and study selection

We conducted a literature search in Medline and Embase to identify original studies and conference abstracts covering the prognostic and predictive value of the CMSs in CRC between 2011 and 2020. No restrictions on study design were set. Key words consisted of ‘colorectal cancer’ and ‘consensus molecular subtypes’. The initial search was conducted in December 2019 and repeated in December 2020.

Eligible studies contained data on the prognostic or predictive value of the CMSs for CRC of any stage. Included studies used either CMS labels or one of the six transcriptomic classifications at the foundation of the CMSs, from which the CMS labels could be deduced (10-15). All methods of CMS labeling were considered eligible. Both study selection based on title and abstract and full text screening were performed by two reviewers independently (SH and TB) using Rayyan software, a web and mobile application for systematic reviews (16). After removal of blinding, the
selected studies were discussed and inclusion of the studies was consensus-based without disagreements between authors.

**Quality assessment and data collection**

The Prediction model Risk Of Bias ASsessment Tool (PROBAST) was used to assess the risk of bias of the included full text studies (17). Two reviewers (SH and TB) completed the tool independently for all studies and the final assessment was consensus-based. No disagreements occurred.

Abstracts, full text manuscripts and supplementary materials were screened for data of interest. Relevant missing data were requested from the authors. Data were extracted by two independent reviewers (SH and TB) and inconsistencies were resolved by consensus. General study characteristics were retrieved, such as study design, GSE accession number of used pre-existing cohorts, number of included patients with CMS label, CRC stage, methods of CMS labeling and distribution of CMSs across the study populations. Outcome data of interest included median absolute OS (mOS) and RFS (mRFS), hazard ratios (HR) for OS, PFS, RFS and SAR from Cox regression analyses and response rates (RR) for systemic therapies, all stratified for CMSs and disease stage. Overlap between studies was accounted for by assessing used cohorts, types of comparisons and analyses performed. When considerable overlap between studies existed, the study with the most robust data on the clinical value was included.

**Statistical analyses**
For the prognostic data, meta-analyses and post-hoc subgroup and sensitivity analyses were performed using the R packages “meta”, “metafor” and “dmetar” (18). For both calculating the pooled mOS times per CMS in mCRC and the CMS comparisons per survival outcome in local and metastatic disease, random-effect inverse variance meta-analyses were performed. The DerSimonian-Laird (DL) \( \tau^2 \)-estimator was used for pooling HRs, whereas the Restricted Maximum-Likelihood (REML) \( \tau^2 \)-estimator was used for pooling median survival times (19). Additional T-tests were performed for pairwise CMS comparisons of pooled mOS times. Data on the predictive value of the CMSs were limited and heterogeneous in terms of treatment strategies, and could therefore not be pooled, but were summarized individually.

Between-study heterogeneity in all meta-analyses was assessed with the Higgins I\(^2\) index and the Cochran’s Q-test(20). Heterogeneity was defined as either an I\(^2\) index \( \geq 50\% \) or a statistically significant Q-test. Sources of heterogeneity were explored by identifying outliers and influential studies (21). Relative contribution of individual studies to heterogeneity compared to their impact on the pooled effect size was visualized with Baujat plots (data not shown) (22). Sensitivity analyses were performed to assess changes in heterogeneity and effect sizes after exclusion of studies that imposed heterogeneity. These studies were not excluded from the final analyses if no statistically significant change of the pooled effects occurred.

To explore the impact on our findings of varying CMS classification methods used in the included papers, we performed subgroup and sensitivity analyses. We compared phenotypic (immunohistochemical [IHC]) classification versus RNA-based classification, single group classification versus direct CMS subtyping and three methods of RNA-based CMS classification. Subgroup analyses were only performed if at least two cohorts with the same classification method could be included per
subgroup; otherwise, sensitivity analyses were run. Subgroup analyses were performed using a random-effect model to pool the effect sizes within the subgroups, and a fixed-effect (plural) model to compare the subgroups (mixed-effect model) (23).

Publication bias was assessed with funnel plots and quantified using the Egger’s test (24,25).

All statistical analyses were performed with R software version 3.6.1 or SPSS version 26. Absolute survival data were presented as medians with 95% confidence interval (95% CI); response data as percentages of the total. Survival comparisons were described with HR, 95% CI and P value and depicted with forest plots. Missing confidence intervals and P values of HRs were calculated using previously described methods (26,27). Distribution of CMS labels in the local versus metastatic setting was assessed with the chi-square statistic (X²) and Pearson residuals. P values were two-sided and a P value of < 0.05 was considered statistically significant. Multiple testing correction, using the Benjamini-Hochberg procedure, was applied to P values from our primary analyses, with an alpha of 0.05. We corrected P values per independent hypothesis, i.e., for the analyses into local CRC, metastatic CRC and mOS per CMS. We used the function ‘mt.rawp2adjp()’ from the R package ‘multtest’ to adjust raw P values according to the classical Benjamini-Hochberg procedure (28).

Results

Selection of eligible studies

The initial search yielded 4,380 studies, of which 67 studies were deemed potentially relevant for further review. After full text screening, 32 additional studies
were excluded. The majority of the exclusions (n=16) were due to considerable overlap in (publicly available) cohorts. In case different studies used similar cohorts, we included those with the most extensive analysis on the prognostic or predictive value of the CMSs (10,13,14,29-41). Other reasons for exclusion after full text screening were no outcome data of interest reported (n=12) (7,42-50), no data on the individual CMSs reported (n=3) (51-53) or a study protocol (n=1) (54) (Figure 1).

Among the 35 eligible studies were 27 full text articles (6,11,12,15,55-77) and eight conference abstracts (78-85). Prognostic data were extracted from 21 studies, of which two were conference abstracts (6,11,15,57,58,60,61,63,64,66-73,75,77,80,81). The predictive value of the CMSs was assessed in sixteen studies and six conference abstracts (12,15,55-57,59,62,64-67,69,72-74,76,78,79,82-85).

Overview of study characteristics and risk of bias

The median size of the study populations with CMS or CMS-related labels was 237 (interquartile range = 113 to 748) (Table 1, Supplementary Table 1 and 2, available online). The majority of the studies was performed in mCRC patients and most studies reported RNA-based CMS labels derived from the primary tumor. OS and RFS were the most frequently reported outcome data. Most predictive studies reported on adjuvant chemotherapy or (the addition of) bevacizumab and cetuximab.

In general, studies were of high quality, as scored with the PROBAST (Supplementary Table 3, available online). Risk of bias was rated as high in five studies, which was commonly due to small sample sizes. Applicability was a concern for fourteen studies, which was invariably related to a custom method of CMS labeling.
or highly specific study populations. When outcome definitions or methods of CMS labeling were lacking, applicability was deemed unclear (n = 3).

CMS distribution

First, we evaluated the relationship between cancer stage and the distribution of reported CMS labels. We observed a statistically significant shift in the CMS distribution between the local and metastatic disease setting ($X^2 = 304.65, P < 0.001$) (Supplementary Table 4, available online). CMS1 was less abundant in the metastatic compared to the localized setting (9.9% versus 18.7%) and conversely an increase in CMS4 tumors was found in metastatic disease (36.1% versus 25.6%). These findings are in line with those in earlier individual studies (6,39).

Prognostic value of the CMSs

**Local disease.** For clinicians and CRC patients, the CMSs could be a valuable tool to estimate disease prognosis. Two studies reported mRFS measures per CMS class for stage II and III patients, predominantly treated with adjuvant chemotherapy, which ranged from 120 to 122 months for CMS1, 114 to 123 months for CMS2, 107 to 115 months for CMS3 and 66 to 120 months in CMS4 (58,72).

The meta-analyses into the prognostic value of the CMSs in local disease included patients from eleven cohorts. We found worse OS for CMS4 compared to CMS1 [stage I-III: HR = 3.28, 95% CI = 1.27 to 8.47; number of included cohorts (n) = 4 (15,61,66,70)] and for CMS4 compared to CMS2 [stage I-III: HR = 2.60, 95% CI = 1.93 to 3.50; n = 3 (11,70)]. The same pattern was observed for the RFS and SAR
To strengthen these findings, we performed two secondary meta-analyses into OS and RFS for stage I-III CMS4 versus the other subtypes, which confirmed the worse survival for CMS4 tumors [OS: HR = 1.67 (95% CI = 1.34 to 2.06), n = 4 (15,31,82,83); RFS: HR = 1.49 (95% CI = 1.11 to 2.00), n = 3 (29,58,60)] (data not graphically shown). Dismal OS and RFS for stage II and III CMS4 tumors were confirmed in four other studies (60,63,71,82), with one study reporting the lowest 5-year lymph node metastasis-free and relapse-free survival for T1 CMS4 tumors (66.7%) (60). This is most probably due to the invasive and mesenchymal properties of CMS4 tumors, resulting in early metastasis (12,86). For CMS3, there is a worse OS compared to CMS2, but better OS than CMS1, although both are not statistically significant. RFS of CMS3 did not differ from CMS1, CMS2 and CMS4 (Supplementary Figure 1, available online) (11,58,61,72,75).

Three (out of ten) conducted pairwise meta-analyses for local disease suffered from between study heterogeneity; differences in sample sizes appeared to explain heterogeneity, as exclusion of these studies reduced heterogeneity without significantly changing the pooled effect sizes (Supplementary Table 5, Supplementary Figure 2, available online).

Metastatic disease. In synchronous or metachronous mCRC, the pooled mOS differed across the CMSs (Figure 3; Supplementary Table 6, available online). Interestingly, CMS1 had the poorest mOS of 11.2 months (6,64,67-70,73), which was significantly worse compared to CMS2 (36.4 months; CMS2 versus CMS1 $P < 0.001$) (6,64,67-70,73,80); CMS3 (21.9 months; CMS3 versus CMS1 $P = 0.02$) (6,64,67-69,73); and CMS4 (26.8 months; CMS4 versus CMS1 $P = 0.001$) (6,64,67-70,73,80). CMS2 had the most favorable OS, which was significantly better than CMS3 ($P = 0.009$) and CMS4 ($P = 0.04$). The mOS of CMS3 and CMS4 were comparable ($P =$
Heterogeneity was present in these meta-analyses and most likely caused by differences in treatment regimens between the cohorts (anti-EGFR treated versus first-line chemotherapy). Exclusion of heterogeneous studies did not significantly alter the pooled survival times of the CMSs (Supplementary Table 7, available online).

Pairwise CMS comparisons in the metastatic setting included patients from nine unique cohorts and showed that CMS1 consistently had worse OS and PFS than CMS2, CMS3 and CMS4 [OS HR range = 0.33 to 0.55; PFS HR range = 0.53 to 0.89 (57,64,67,69,70,81)] (Figure 2, A and C, Supplementary Figure 1A, available online). This is in line with the reported mOS times. The mesenchymal CMS4 subtype was associated with worse OS compared to the epithelial CMS2 subtype [HR = 1.41, 95% CI = 1.07 to 1.86; n = 6 (57,69,70,77)], but PFS did not differ (Figure 2B) (57,69,70). CMS3 did not show statistically significant differences in OS and PFS compared to CMS2 and CMS4, although PFS tended to be worse for CMS3 when compared to CMS4 (57,69) (Supplementary Figure 1, available online). Heterogeneity was detected in five out of twelve meta-analyses, which could partly be explained by differences in cohort sizes, and exclusion of these studies did not change the summary effect sizes (Supplementary Table 5, available online).

Predictive value of the CMSs

Local disease. Currently, patient selection for adjuvant chemotherapy is suboptimal, leading to over- and undertreatment in a large proportion of patients (2). Differences in tumor biology, as captured by the CMSs, could serve as predictors for treatment benefit. As data on the predictive value of the CMSs were rather heterogeneous in terms of received treatment regimens, we summarized the results
of each study individually. Adjuvant chemotherapy, as compared to surgery alone, improved OS for both stage II and III in CMS2 and CMS3 (HR range = 0.16 to 0.45) (Figure 4A) (15,55,66). Adjuvant chemotherapy was not clearly associated with better survival in CMS1 and CMS4 tumors (Figure 4, A-B) (59). With respect to combination regimens for adjuvant treatment, the addition of oxaliplatin to 5-FU increased RFS in stage III CMS2 patients in one study (HR = 0.61, 95% CI = 0.43 to 0.87) (Figure 4C) (72), whereas in another study CMS4 patients seemed not to benefit from this regimen (83). The difference in treatment efficacy between CMS2/CMS3 and CMS1/CMS4 likely results from intrinsic molecular differences between epithelial versus mesenchymal and immunogenic tumors.

Metastatic disease. All predictive studies in mCRC assessed first-line treatment regimens, except for two studies and one conference abstract reporting on more recently developed targeted therapies in the second-line (62,65,84). Strikingly, all studies uniformly showed that an irinotecan backbone increased both PFS and OS in patients with CMS4 tumors (PFS HR range = 0.31 to 0.72; OS HR range = 0.43 to 0.45) (Figure 5A; Supplementary Figure 3A, available online) (57,69,78,85). One abstract reported an increased OS for oxaliplatin- compared to irinotecan-based treatment in CMS4 patients (HR = 0.42, 95% CI = 0.17 to 1.00) (79). However, irinotecan-based therapy was given to a minority of patients compared to oxaliplatin-based chemotherapy (19.7% versus 80.3%), which might have influenced the results. For CMS1, a slight increase of survival with an irinotecan-based regimen was also observed (Figure 5A; Supplementary Figure 3A, available online). These findings are of importance considering that the first-choice regimen in the first-line setting often is oxaliplatin-based and, consequently, a large group of CMS1 and CMS4 mCRC patients may receive suboptimal first-line treatment.
Regarding targeted therapies, a PFS benefit from adding bevacizumab to capecitabine in both CMS2 and CMS3 was seen (HR = 0.44 [95% CI = 0.29 to 0.68] and HR = 0.35 [95% CI = 0.14 to 0.86], respectively) (Supplementary Figure 3B, available online). For OS, the same trend was observed (Figure 5B) (67). Adding bevacizumab to FOLFIRI (but not to capecitabine), as compared to FOLFIRI monotherapy, was associated with better OS (HR = 0.38 95% CI = 0.14 to 1.06) and PFS in CMS4 (Figure 5A; Supplementary Figure 3B, available online) (57,67). However, when comparing FOLFIRI combined with either bevacizumab or cetuximab in CMS4 patients, bevacizumab was associated with inferior outcomes compared to cetuximab for both OS (HR = 1.75, 95% CI = 1.16 to 2.63) and PFS (Figure 5C; Supplementary Figure 3C, available online) (73). The combination of bevacizumab with FOLFOX appeared more effective than cetuximab plus FOLFOX in CMS1 for both OS (HR = 0.43, 95% CI = 0.27 to 0.68) and PFS (Figure 5C; Supplementary Figure 3C, available online) (64).

The addition of cetuximab to either FOLFOX or CAPOX-bevacizumab (CAPOX-B) was beneficial in KRAS wildtype (KRAS wt) CMS2 patients for OS (FOLFOX: HR = 1.61, 95% CI = 1.16 to 2.22; CAPOX-B: HR = 1.92, 95% CI = 1.15 to 3.23) and PFS (Figure 5C; Supplementary Figure 3C, available online) (64,74). One abstract also showed an increased OS for anti-EGFR in CMS2 and 3 compared to bevacizumab (HR = 0.67, 95% CI = 0.42 to 1.06) (79). Furthermore, evidence for anti-EGFR benefit in CMS2 was observed in terms of RECIST-defined tumor response rates, showing a very high disease control rate for both CMS2 (92.3%) and CMS4 (81.8%) with anti-EGFR therapy (+/- 80.0% combined with irinotecan and +/- 20.0% monotherapy) (56,69,76). Additionally, clinical response and disease-free survival of cetuximab monotherapy were found to be better for KRAS wt CMS2 patients (12).
Two articles and one conference abstract reported on targeted therapies as second-line regimens for mCRC. One study showed an increase in PFS when CMS4 tumors were treated with nintedanib (triple angiokinase inhibitor) as compared to placebo (HR = 0.60, 95% CI = 0.40 to 0.91); no effect was seen in CMS2 tumors. Unclassified patients had both an increased OS (HR = 0.29, 95% CI = 0.14 to 0.52) and PFS (HR = 0.20, 95% CI = 0.08 to 0.49) (65). The combination of napabucasin (blocks STAT3; downregulates IDO1 and PD-L1) and pembrolizumab (anti-PD1) showed a response rate of 33.0% for CMS1 and CMS4 tumors (62). The third study showed no clear benefit for CMS4 from bintrafusp alfa, targeting both TGF-β and PD-L1, with or without radiotherapy (84).

Altogether, these data indicate that CMS2 and CMS3 could predict benefit from adjuvant chemotherapy. In mCRC, CMS4 has predictive value for the use of first-line irinotecan-based chemotherapy. Also, combining an irinotecan backbone with bevacizumab or cetuximab (for KRAS wt patients) is favorable in CMS4. CMS1 appears to be predictive for the addition of bevacizumab and CMS2 has potential predictive value for the use of cetuximab in KRAS wt patients. Immunotherapy might be an effective treatment for CMS1 mCRC.

**Impact of CMS classification methods**

Differences in CMS classification methods could impact the estimated prognostic and predictive values of the individual CMSs as determined in our study. The majority of included studies used gene expression data and the CMS R package (n = 15). Other classification methods were NanoString-based (n = 7), IHC-based (n
= 4), the CMScaller (n = 1), the multinomial elastic net classifier (n = 1) or single group classification from which the CMSs could be deduced (n = 4).

We previously described a shift in distribution of CMS labels between the local and metastatic setting and this remained the same when excluding those studies using classification methods with only three CMS classes (IHC, Roepman and De Sousa subtypes) [Supplementary Table 8, available online]. Studies in which the distribution of CMS labels was markedly altered are indicated in Supplementary Table 1 and 2. Most differences were due to the usage of the IHC CMS classifier, as this cannot detect CMS3 labels, or due to a specific subgroup of patients (e.g., T1 tumors or a RAS wildtype cohort). Furthermore, it was recognized that specific classification methods are associated with slightly different CMS distributions [Supplementary Figure 4, available online]. In the local setting, more CMS3 and unknown labels were observed when using the CMS R package, but reduced CMS4 labels. The IHC method yielded more CMS2 and reduced unknown labels. In the metastatic setting, NanoString classifiers detected less CMS3, and the CMS R package less CMS2 and CMS4 labels.

To explore the effect of the different classification methods on the prognostic value of the CMSs, we looked into three comparisons: transcriptomic versus IHC-based labeling, single group transcriptomic labels as surrogates for CMS labels versus direct CMS labeling and CMS R package classification versus other RNA-based methods. Two analyses included IHC-based CMS labels next to RNA-based CMS labels [Supplementary Tables 9 and 10, available online]. In sensitivity analyses, no major changes in (direction of) the pooled effect sizes occurred after exclusion of IHC-based data [Supplementary Table 10, available online]. Nine meta-analyses included single group transcriptomic labels (Budinska, Roepman and Sadanandam)
as surrogates for CMS labels. For both Budinska and Roepman labels, neither subgroup nor sensitivity analyses showed a statistically significant difference in prognostic value compared to CMS labels (Supplementary Table 11 and 12, available online). Two subgroup analyses for Sadanandam labels also showed similar prognostic value compared to CMS classes, but a third subgroup analysis found a statistically significant difference between these taxonomies (CMS4 versus CMS1, RFS local) (Supplementary Table 11, available online). However, the direction of both effect sizes was similar and therefore our conclusions were not affected. Furthermore, subgroup and sensitivity analyses showed no statistically significant differences between the different RNA-based CMS classification methods, except for one subgroup analysis into NanoString-based classification versus the CMS R package (CMS2 mOS) (Supplementary Table 13, 14 and 15, available online). This again had no effect on the interpretation of our findings.

Regarding the predictive value, three different CMS classification methods were used in the studies investigating localized disease [the CMS R package (55), single group classification (15) and the IHC-classifier (66)]. The OS benefit of adjuvant chemotherapy compared to surgery alone for CMS2 and CMS3 was consistent across the four cohorts, despite using different classifiers. As for stage II CMS4, no clear benefit from adjuvant chemotherapy was detected by using both the CMS R package and the IHC-classifier (55,59). In the metastatic setting, the beneficial effect of irinotecan-based chemotherapy for CMS4 as compared to oxaliplatin-based chemotherapy was mainly based on CMS labels from the CMS R package (three studies), although the same effect was found with a NanoString classifier (78). The favorable effect of cetuximab combined with an oxaliplatin-based backbone for CMS2 patients was shown with both NanoString-based and IHC-based CMS classification
(64,74). Other supporting evidence for the benefit of anti-EGFR therapy in CMS2 patients was shown with the IHC-classifier, as compared to the addition of bevacizumab (79), but also with NanoString-based classification (56), the CMS R package (69) and the CMScaller (76).

Altogether, it seems that the prognostic and predictive value of the CMSs are robust and remain unaffected by the classification method employed. The possibility to use a variety of CMS classification methods with largely similar CMS detection and related clinical associations could be a major advantage for future implementation of the CMSs in the clinic.

**Publication bias**

In the majority of 21 out of 26 meta-analyses, publication bias did not play a statistically significant role, as assessed in funnel plots and quantified with the Egger’s test (Supplementary Figures 5-7, available online). In five meta-analyses into mCRC, i.e., OS and PFS of CMS4 versus CMS1, OS of CMS2 versus CMS1, OS of CMS3 versus CMS1 and absolute mOS of CMS1, publication bias could have influenced our outcomes, warranting careful interpretation of these specific results. However, in general the role of publication bias in our meta-analyses seemed limited.

**Discussion**

To our knowledge, this is the first systematic review and meta-analysis providing a comprehensive overview of the prognostic and predictive significance of the CMS classification (6). Our study underlines the additive value of molecular CMS
classification and its potential for patient stratification in clinical trials and future clinical use, both with regards to estimating disease prognosis and selecting patients for specific treatment regimens. These benefits seem not to be restricted to a specific classification method, but appear to be rather universal for a variety of molecular classifiers, depending on the materials and resources available.

We demonstrated with aggregated data the unfavorable prognostic value of CMS4 in local disease compared to CMS1 and CMS2. As an example, one of the included studies looked into a relatively large cohort of 860 stage III patients and identified that CMS4 performed significantly worse than CMS1 and CMS2 with regards to OS (70). In mCRC, we found that CMS1 was associated with worse survival compared to the other subtypes. In comparison with our pooled data, the largest included study in the metastatic setting investigated a cohort of 581 patients and found an mOS of fifteen months for CMS1, and a significantly worse outcome compared to CMS2 and CMS4 (64). This is in line with published data showing worse survival for metastasized BRAF mutated and MSI tumors (87-89). As both MSI and BRAF mutations are typical properties of the CMS1 subtype, they might explain the worse survival of metastatic CMS1 tumors, as identified in our study. In both local and metastatic disease, we identified CMS2 as the subtype with the most favorable prognosis.

Our findings provide insight into the variable efficacy of systemic therapies in CRC, currently leading to substantial over- and undertreatment in a large proportion of patients. In local disease, CMS2 and CMS3 predict benefit from adjuvant chemotherapy, in contrast to the mesenchymal and immune subtypes. The lack of fluoropyrimidine-monotherapy benefit for dMMR CMS1 tumors could be explained by the fact that a sufficient MMR machinery is required for FU-induced cell death (90-92).
CMS4 tumors possibly behave more like the early-dissemination model instead of the classical linear-progression model (93,94). Therefore, a large proportion of patients diagnosed with local disease might already have (yet undetectable) micro-metastases and, consequently, lacking benefit from adjuvant chemotherapy. The predictive relevance of the CMSs for neoadjuvant chemotherapy in resectable stage II and III colon cancer will be further evaluated in the CONNECTION II trial (95).

In CMS4 metastatic disease, studies showed better efficacy of an irinotecan-based first-line regimen as compared to an oxaliplatin-containing regimen. One of these is the recently published retrospective analysis of the phase III TRIBE2 study on the superiority of bevacizumab combined with triplet chemotherapy (FOLFOXIRI) as compared to doublets. Adding irinotecan to doublet therapy with bevacizumab showed a statistically significant benefit for CMS4 tumors as compared to doublet therapy with bevacizumab (78). The phase III TRICOLORE trial also found the best effectiveness of an irinotecan-based regimen (S-1/irinotecan/bevacizumab) for CMS4 compared to an oxaliplatin backbone combined with bevacizumab (85). Currently, standard first-line regimens in mCRC are often oxaliplatin-based, which stresses the relevance of stratifying patients according to molecular features to prevent suboptimal treatment. In addition, the majority of peritoneal metastases are of the CMS4 subtype. Therefore, intraperitoneal irinotecan-based treatment holds great promise in light of our findings and will be assessed in the phase I INTERACT trial (96,97). CMS4 mCRC also predicted benefit from irinotecan combined with cetuximab (in KRAS wt tumors) and bevacizumab (in KRAS mutant tumors) and thus, combinations of these targeted agents with an irinotecan backbone could serve as preferred first-line options in CMS4 mCRC patients.
CMS1 tumors likely benefit most from immunotherapy in metastatic disease. The available studies using CMS1 instead of MSI are sparse, but based on the clear efficacy in these studies we assume that immunotherapy will be the treatment of choice (4, 98, 99). As second-line regimen for CMS1 tumors, a chemotherapy backbone combined with bevacizumab could be an effective alternative, as was shown in the retrospective analysis of the phase III CALGB/SWOG trial, that assessed the relative efficacy of cetuximab versus bevacizumab when added to standard FOLFOX or FOLFIRI-based first-line chemotherapy (Figure 5C) (64). The preference for bevacizumab rather than cetuximab in CMS1 could be explained by the important role of T-lymphocytes (abundant in CMS1 tumors) in vessel normalization induced by blockade of angiogenesis. This results in reduced tumor hypoxia and vessel leakage, increasing influx of immune effector cells and chemotherapy into the tumor parenchyma (7, 100, 101).

Cetuximab benefit was most evident in patients with CMS2 KRAS wt mCRC. It was previously shown that CMS2 tumors are predominantly left-sided and specifically these tumors are characterized by an activated EGFR pathway, rendering them susceptible for cetuximab treatment (64, 102). Previous studies have also shown that combined APC and TP53 mutations are related to cetuximab sensitivity, and these mutations are enriched in CMS2 patients (75.0% versus 17.0 to 41.0% in CMS1-4) (103, 104).

An important and ongoing topic of discussion is the additive value of the CMSs as compared to standardly employed TNM staging, MMR and gene mutation status. With regards to the latter, a strong CMS-dependent prognostic impact of BRAFV600E and KRAS mutations in MSS tumors has been demonstrated, with the strongest prognostic value of BRAFV600E mutations in CMS1 tumors and KRAS mutations in
CMS2 and CMS3 tumors (105). With respect to TNM staging, we recently reported that gene signatures of advanced-stage CRC were highly correlated with CMS4 and identified a CMS4-subgroup within high-risk stage II patients with worse survival than the total group of stage II patients, stressing the additive prognostic value of transcriptomic subtypes to TNM-staging (39). Furthermore, a retrospective study in early-stage CRC demonstrated that tumor microenvironment features, such as cancer-associated-fibroblasts (CAFs) and CD3+ and CD8+ T cell infiltration, are particularly prognostic as compared to T- and N-stage, MMR-, BRAF- and KRAS mutational status, and CMS4 score (58). As microenvironment features, including CAFs and immune cell infiltration, are intrinsic properties of the CMSs, this suggests that transcriptomic profiling might be more accurate in prognosis prediction as compared to TNM staging. This, however, requires further refinement into more specific prognostic subgroups, that go beyond sole CMS labeling. For example, next to CAF and CytoLym scores, several studies within the established CMSs have shown the favorable prognostic value of immune markers, such as expression of chemokine-like factor in CMS1 (59), immunoscore-like metagenes within CMS2 and CMS3 (36), and both specific transcriptomic immune subtypes and immune checkpoint metagenes in CMS3 (36,106). For CMS1 and CMS3, deregulated Bcl2-dependent apoptosis was a poor prognostic marker (107), whereas in CMS4 factors related to the anti-oxidant response and oxidative DNA damage conferred poor prognosis (108).

As for predicting treatment benefit, the additive value of transcriptomic subtyping to current standard of care is promising, providing a full overview of molecular targets and exploitable tumor vulnerabilities for each tumor, enabling better selection for commonly used chemotherapeutic regimens and bringing newly discovered targeted agents within reach for each patient. For example, in preclinical
models, CMS1 and CMS4 predicted response to heat shock protein 90 (HSP90) inhibitors (31), but also, response to triple therapy with trametinib (MEK1/2-inhibitor), neratinib (pan-ERBB inhibitor) and trastuzumab (HER2-inhibitor) (109). On the other hand, CMS2 and CMS3 were more sensitive to aurora kinase inhibitors (10). These studies point out differences in treatment sensitivity between CMS1 and CMS4 (mesenchymal/immune) versus CMS2 and CMS3 (epithelial) tumors.

So far, clinical implementation of molecular classification systems has proven to be challenging. Extensive genomic and transcriptomic profiling of tumors is time-consuming and costly, and its cost-effectiveness has yet to be demonstrated. Furthermore, RNA from formalin-fixed paraffin-embedded (FFPE) tissue, which is widely used in hospitals, is vulnerable and generally of low quality. The CMS classification must be optimized for use on this material and a standardized pipeline needs to be developed to determine individual patients’ subtypes while accounting for intra-tumor heterogeneity. We showed that the value of CMS classification is not restricted to a specific classification method, but appears to be rather universal for a variety of molecular classifiers, which could enhance clinical implementation. Further advancements in this field have been made, as Morris et al. most recently developed a 100-gene FFPE NanoString classifier with high classification accuracy (>80.0% compared to the original CMS classifier) and validated it in a CLIA-certified laboratory (68). Similarly, for Sadanandam subtypes, a 38-gene NanoString CRCA classifier (NanoCRCA) was developed, that showed high concordance between classification of FFPE- and FF-material, and with CMS subtypes (75.0%) (110). Alternatively, image-based transcriptomic classification could be further explored, as this circumvents abovementioned issues and seems promising in first attempts (111).
Additional studies that address these issues are warranted to pursue successful implementation of molecular classification systems in the clinic.

The limitations of our systematic review and meta-analysis arise mainly from the retrospective nature of the included studies. Survival data were either incomplete or lacking and therefore median survival per CMS in the curative setting and some pairwise CMS comparisons per survival outcome could not be pooled. Number of studies included in the meta-analyses of localized CRC were limited and more often had an applicability concern in our risk of bias assessment, which warrants careful interpretation of these results. Using data from retrospective cohorts with differences in samples sizes, CMS classification methods and treatments, imposed heterogeneity on our meta-analyses. Most pooled cohorts used for the prognostic meta-analyses included patients that were treated with adjuvant chemotherapy in the curative setting and first-line chemotherapy in the metastatic setting. However, in the curative setting we could not completely rule out differences in received treatment, which has affected the pooled estimates. However, sensitivity analyses showed no statistically significant change of effect sizes after exclusion of cohorts that imposed heterogeneity, indicating robustness of our meta-analyses (Supplementary Table 5, available online). Due to limited data on the predictive value of the CMSs, a traditional meta-analysis was not feasible. Consequently, although insightful, the results of the predictive potential of the CMSs are preliminary and these findings should be validated in future randomized trials. Lastly, publication bias, as observed in funnel plots and quantified with the Egger's test, was a minor concern for the majority of our meta-analyses (Supplementary Figures 5-7, available online).

Importantly, the varying CMS classification methods used in the included studies have an influence on the CMS distribution of patient cohorts. Labeling of
individual samples can vary by methods of normalization of the raw data, settings of the specific classifier (e.g., SSP versus RF in CMS R package) and preselection of specific patient groups. Therefore, we thoroughly explored the effect of different CMS labeling methods on our pooled prognostic outcomes, and we could not identify any difference between the CMS classifiers that would change our findings. The same holds true for the summarized predictive value of the CMSs. Although CMS labeling varies to some extent between the classifiers, it seems that the subtypes are sufficiently different to be detected, regardless of classification method and preselection of cohorts.

This comprehensive systematic review and meta-analysis of the CMS classification demonstrates that it is a robust framework for both prognostication and predicting treatment benefit in a biology-driven patient-centered fashion. Future studies should include molecular subtyping in their trial design to validate these associations. In addition, refinement of the subtype classification using prognostic biomarkers may lead to an even more thorough understanding of the different subtypes. Once the CMSs are sufficiently validated in prospective clinical trials, in addition to TNM staging, they could have great impact on improving and tailoring treatment strategies for the individual CRC patient.

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Author contributions: SH, TB, DS and LV designed the study. SH and TB identified eligible studies, assessed quality of the included studies, extracted data and performed data analyses. SH, TB, DS and LV wrote and revised the manuscript.

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Data Availability

Data used in this systematic review and meta-analysis could be provided upon request, only with permission of the authors of the original studies.
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### Table 1. Overview of study characteristics

| Characteristic                        | Studies, No. (%) | Prognostic, No. (%) | Predictive, No. (%) |
|---------------------------------------|------------------|---------------------|---------------------|
| Total No.                             | 35               | 21                  | 22                  |
| Papers                                |                  |                     |                     |
| Full text articles                    | 27 (77.1)        | 19 (90.5)           | 16 (72.7)           |
| Abstracts                             | 8 (22.9)         | 2 (9.5)             | 6 (27.3)            |
| Study design                          |                  |                     |                     |
| Retrospective                         | 31 (88.6)        | 20 (95.2)           | 19 (86.4)           |
| Prospective                           | 4 (11.4)         | 1 (4.8)             | 3 (13.6)            |
| Disease stages in survival analysis   |                  |                     |                     |
| II and III                            | 7 (20.0)         | 5 (23.8)            | 4 (18.2)            |
| III                                   | 3 (8.6)          | 1 (4.8)             | 2 (9.1)             |
| IV                                    | 17 (48.6)        | 9 (42.9)            | 13 (59.1)           |
| I-IV                                  | 3 (8.6)          | 1 (4.8)             | 2 (9.1)             |
| Other                                 | 5 (14.3)         | 5 (23.8)            | 1 (4.5)             |
| Transcriptomic classification          |                  |                     |                     |
| Consensus molecular subtypes (CMS)    |                  |                     |                     |
| CMS R package                         | 16 (45.7)        | 11 (52.4)           | 9 (40.9)            |
| NanoString                            | 8 (22.6)         | 4 (19.0)            | 5 (22.7)            |
| Immunohistochemistry                  | 4 (11.4)         | 2 (9.5)             | 3 (13.6)            |
| Other                                 | 2 (5.7)          | 1 (4.8)             | 1 (4.5)             |
| Sadanandam subtypes                   | 2 (5.7)          | 1 (4.8)             | 2 (9.1)             |
| De Sousa e Melo subtypes              | 1 (2.9)          | 0 (0.0)             | 1 (4.5)             |
| Budinska subtypes                     | 1 (2.9)          | 1 (4.8)             | 0 (0.0)             |
| Roepman subtypes                      | 1 (2.9)          | 1 (4.8)             | 1 (4.5)             |
| Survival outcomes                     |                  |                     |                     |
| Overall survival                      | 26 (74.3)        | 16 (76.2)           | 17 (77.3)           |
| Relapse-free survival                 | 13 (37.1)        | 10 (47.6)           | 5 (22.7)            |
| Progression-free survival             | 12 (34.3)        | 7 (33.3)            | 10 (45.5)           |
| Survival after relapse                | 3 (8.6)          | 3 (14.3)            | 0 (0.0)             |
| Systemic treatment                    |                  |                     |                     |
| Adjuvant chemotherapy                 | 6 (17.1)         | 3 (14.3)            | 6 (27.3)            |
| Targeted agents                       | 11 (31.4)        | 4 (19.0)            | 11 (50.0)           |
| First-line chemotherapy metastatic colorectal cancer | 5 (14.3) | 2 (9.5) | 5 (22.7) |
| Immunotherapy                         | 2 (5.7)          | 0 (0.0)             | 2 (9.1)             |
| No. of classified samples             |                  |                     |                     |
| Median                                | 237              | 237                 | 273                 |
| Interquartile range                   | 113-748          | 113-765             | 168-660             |

*a The percentages were calculated using the total number in the first row for each column.
**Figures**

**Figure 1.** Flowchart of study identification and selection process. a Eight studies were used for both prognostic and predictive analysis.

**Figure 2.** Forest plots of pooled and single hazard ratios (HRs) for different survival outcomes per pairwise consensus molecular subtype (CMS) comparison in colorectal cancer. (A) CMS4 versus CMS1. (B) CMS4 versus CMS2. (C) CMS2 versus CMS1. Number of cohorts (n) and total number of included patients (N) per meta-analysis indicated with n/N; as several studies described more cohorts, the number of cohorts n does not reflect the number of studies. adjP indicates significance for the random-effect model (corrected for multiple testing). Heterogeneity depicted as I² index and Cochran’s Q-test P value. All tests were 2-sided. a Estimate. CI = confidence interval; HR = hazard ratio; adjP = adjusted P value; NA = not applicable.

**Figure 3.** Median overall survival in metastatic disease per consensus molecular subtype (CMS). Vertical lines indicate statistically significant comparisons (T-test). aP < 0.001; bP = 0.02; cP = 0.001; dP = 0.009; eP = 0.04 (corrected for multiple testing). Number of cohorts (n) and total number of included patients (N) per meta-analysis indicated with n/N; as several studies described more cohorts, the number of cohorts n does not reflect the number of studies. Heterogeneity depicted as I² index and Cochran’s Q-test P value. All tests were 2-sided.
**Figure 4.** Forest plots of hazard ratios (HRs) for the predictive value of the consensus molecular subtypes (CMSs) for adjuvant chemotherapy in colorectal cancer. (A) Overall survival for adjuvant chemotherapy versus surgery alone. (B) Relapse-free survival for adjuvant chemotherapy versus surgery alone. (C) Relapse-free survival for oxaliplatin plus 5-fluorouracil and leucovorin versus 5-fluorouracil plus leucovorin.

*a*Immunohistochemistry classification: CMS2/3 depicted as CMS2. 5-FU = 5-fluorouracil; FOLFOX = 5-fluorouracil, leucovorin, and oxaliplatin; FULV = 5-fluorouracil and leucovorin.

**Figure 5.** Forest plots of hazard ratios (HRs) for the predictive value of the consensus molecular subtypes (CMSs) for first-line systemic therapy in metastatic colorectal cancer. (A) Progression-free survival for an irinotecan backbone versus control. (B) Overall survival for the addition of bevacizumab versus control. (C) Overall survival for the addition of bevacizumab versus cetuximab. 

*a*Total number of patients, no information on the number of CMS4 patients. 

*b*KRAS wildtype population. 

*c*75.2% FOLFOX / 24.8% FOLFIRI. 

*d*Immunohistochemistry classification: CMS2/3 depicted as CMS2. Bev = bevacizumab; Cap = capecitabine; CAPOX = capecitabine and oxaliplatin; Cet = cetuximab; CM = classification method; FOLFIRI = 5-fluorouracil, leucovorin, and irinotecan; FOLFOX = 5-fluorouracil, leucovorin, and oxaliplatin; mFOLFOX = modified FOLFOX; FOLFOXIRI, 5-fluorouracil, leucovorin, oxaliplatin and irinotecan; IRI = irinotecan; Mit = mitomycin; N = number of patients.
Records retrieved from Medline (n = 2328)
Records retrieved from Embase (n = 3818)
Additional records from other sources (n = 10)

Total number of records from searches (n = 6156)

Duplicates removed (n = 1776)

Screened by title and abstract (n = 4380)

Excluded for full-text screening (n = 4313)
Not including humans
No consensus molecular subtype labels
Patients with other diseases
Reviews

Full-text articles assessed for eligibility (n = 67)

Studies excluded (n = 32)
No new dataset or analysis (n = 16)
No outcome data (n = 12)
No individual CMS data (n = 3)
Study protocol (n = 1)

Studies included (n = 35)

Studies included in prognostic analysis (n = 21)
Studies included in predictive analysis (n = 22)
## A

| Survival                        | n/N    | HR   | 95% CI          | CMS4 versus CMS1 | adjP | Heterogeneity |
|---------------------------------|--------|------|-----------------|------------------|------|---------------|
| Overall survival                |        |      |                 |                  |      |               |
| Stage I–III                     | 4/1327a| 3.28 | (1.27 to 8.47)  |                  | 0.03 | I² = 83.0%, P < 0.001 |
| Stage IV                        | 7/601  | 0.45 | (0.32 to 0.62)  |                  | <0.001 | I² = 49.7%, P = 0.06 |
| Progression-free survival       |        |      |                 |                  |      |               |
| Stage IV                        | 7/538  | 0.53 | (0.38 to 0.75)  |                  | <0.001 | I² = 58.6%, P = 0.047 |
| Relapse-free survival           |        |      |                 |                  |      |               |
| Stage II–III                    | 5/2410 | 1.84 | (1.35 to 2.51)  |                  | <0.001 | I² = 68.4%, P = 0.01 |
| Survival after relapse          |        |      |                 |                  |      |               |
| Stage III                       | 1/78   | 1.11 | (0.59 to 2.50)  |                  | 0.78 | NA            |

Meta-analysis                    | Single study  
|---------------------------------|----------------|

favors CMS4  favors CMS1

## B

| Survival                        | n/N    | HR   | 95% CI          | CMS4 versus CMS2 | adjP | Heterogeneity |
|---------------------------------|--------|------|-----------------|------------------|------|---------------|
| Overall survival                |        |      |                 |                  |      |               |
| Stage I–III                     | 3/1337 | 2.60 | (1.93 to 3.50)  |                  | <0.001 | I² = 00.0%, P = 0.38 |
| Stage IV                        | 6/414  | 1.41 | (1.07 to 1.86)  |                  | 0.04 | I² = 29.3%, P = 0.22 |
| Progression-free survival       |        |      |                 |                  |      |               |
| Stage IV                        | 5/376  | 0.96 | (0.72 to 1.29)  |                  | 0.80 | I² = 39.3%, P = 0.16 |
| Relapse-free survival           |        |      |                 |                  |      |               |
| Stage II–III                    | 8/4117 | 1.46 | (1.29 to 1.64)  |                  | <0.001 | I² = 00.0%, P = 0.76 |
| Survival after relapse          |        |      |                 |                  |      |               |
| Stage II–III                    | 3/334  | 1.97 | (1.41 to 2.74)  |                  | <0.001 | I² = 00.0%, P = 0.99 |

Meta-analysis                    | Single study  
|---------------------------------|----------------|

favors CMS4  favors CMS2

## C

| Survival                        | n/N    | HR   | 95% CI          | CMS2 versus CMS1 | adjP | Heterogeneity |
|---------------------------------|--------|------|-----------------|------------------|------|---------------|
| Overall survival                |        |      |                 |                  |      |               |
| Stage I–III                     | 3/1771a| 0.92 | (0.70 to 1.21)  |                  | 0.74 | I² = 00.0%, P = 0.52 |
| Stage IV                        | 8/840  | 0.33 | (0.23 to 0.48)  |                  | <0.001 | I² = 62.8%, P = 0.009 |
| Progression-free survival       |        |      |                 |                  |      |               |
| Stage IV                        | 7/684  | 0.53 | (0.35 to 0.81)  |                  | 0.01 | I² = 74.7%, P < 0.001 |
| Relapse-free survival           |        |      |                 |                  |      |               |
| Stage II–III                    | 4/2723 | 0.97 | (0.82 to 1.15)  |                  | 0.74 | I² = 00.0%, P = 0.84 |
| Survival after relapse          |        |      |                 |                  |      |               |
| Stage II–III                    | 1/92   | 0.27 | (0.12 to 0.62)  |                  | 0.002 | NA            |

Meta-analysis                    | Single study  
|---------------------------------|----------------|

favors CMS2  favors CMS1
CMS1 (n/N = 9/289) \( I^2 = 70.3\%, P < 0.001 \)
CMS2 (n/N = 9/858) \( I^2 = 87.7\%, P < 0.001 \)
CMS3 (n/N = 7/292) \( I^2 = 84.6\%, P < 0.001 \)
CMS4 (n/N = 10/770) \( I^2 = 91.7\%, P < 0.001 \)
## A

| Treatment          | Control       | Stage | N   | HR    | 95% CI       | Overall survival |
|--------------------|---------------|-------|-----|-------|--------------|------------------|
| 5-FU based         | Surgery       | II    | 59  | 1.10  | (0.32 to 4.10) |                 |
| Allen et al. 2018  |               |       |     | 1.19  | (0.31 to 4.06) |                 |
| CM: CMS R package  |               |       |     | 1.02  | (0.29 to 3.67) |                 |
| Chemotherapy       | Surgery       | II    | 30  | 0.65  | (0.04 to 11.0) |                 |
| Li et al. 2020†    |               |       |     | 0.76  | (0.05 to 12.1) |                 |
| CM: Immunohistochemistry |            |       |     | 0.76  | (0.05 to 12.1) |                 |
| FOLF                | Surgery       | III   | 41  | 0.49  | (0.15 to 1.60) |                 |
| Allen et al. 2018  |               |       |     | 0.49  | (0.15 to 1.60) |                 |
| CM: CMS R package  |               |       |     | 0.49  | (0.15 to 1.60) |                 |
| 5-FU                | Surgery       | III   | 46  | 0.39  | (0.10 to 1.54) |                 |
| Roepman et al. 2014|               |       |     | 0.39  | (0.10 to 1.54) |                 |
| CM: CMS Related    |               |       |     | 0.39  | (0.10 to 1.54) |                 |

## B

| Treatment          | Control       | Stage | N   | HR    | 95% CI       | Relapse-free survival |
|--------------------|---------------|-------|-----|-------|--------------|------------------------|
| FULV               | Surgery       | II    | 48  | 3.64  | (0.87 to 15.3) |                        |
| Dunne et al. 2016  |               |       |     | 3.50  | (0.87 to 15.3) |                        |
| CM: CMS R package  |               |       |     | 1.45  | (0.64 to 3.32) |                        |
| Chemotherapy       | Surgery       | II    | 30  | 2.22  | (0.55 to 9.03) |                        |
| Li et al. 2020†    |               |       |     | 2.66  | (0.55 to 9.03) |                        |
| CM: Immunohistochemistry |            |       |     | 0.81  | (0.49 to 1.33) |                        |
| FULV               | Surgery       | III   | 32  | 1.77  | (0.40 to 7.94) |                        |
| Dunne et al. 2016  |               |       |     | 1.77  | (0.40 to 7.94) |                        |
| CM: CMS R package  |               |       |     | 0.60  | (0.27 to 1.61) |                        |

## C

| Treatment          | Control | Stage | N   | HR    | 95% CI       | Relapse-free survival |
|--------------------|---------|-------|-----|-------|--------------|------------------------|
| FOLFOX             | FULV    | II    | 141 | 0.76  | (0.21 to 2.84) |                        |
| Song et al. 2016   | FULV    |       |     | 0.76  | (0.21 to 2.84) |                        |
| CM: CMS Related    | FULV    |       |     | 0.75  | (0.23 to 2.83) |                        |
| FOLFOX             | FULV    | III   | 231 | 0.77  | (0.46 to 1.26) |                        |
| Song et al. 2016   | FULV    |       |     | 0.77  | (0.46 to 1.26) |                        |
| CM: CMS Related    | FULV    |       |     | 0.55  | (0.24 to 1.60) |                        |

---

**Figure 4**
### A

| Treatment                  | Control           | N   | HR   | 95% CI       | Progression-free survival |
|----------------------------|-------------------|-----|------|--------------|---------------------------|
| Irinotecan–based           | Oxaliplatin–based | 21  | 0.64 | (0.20 to 1.69) |                           |
| Okita et al. 2018          |                   | 47  | 1.09 | (0.55 to 2.06) |                           |
| CM: CMS R package          |                   | 63  | 0.67 | (0.37 to 1.15) |                           |
|                            |                   | 48  | 0.31 | (0.13 to 0.64) |                           |
| FOLFIRI                    | FOLFOX            | 10  | 1.43 | (0.28 to 7.28) |                           |
| Del Rio et al. 2017        |                   | 24  | 1.17 | (0.50 to 2.75) |                           |
| CM: CMS R package          |                   | 21  | 0.66 | (0.24 to 1.82) |                           |
|                            |                   | 27  | 0.35 | (0.12 to 1.04) |                           |
| S-1+IRI+Bev                | mFOLFOX+Bev/CAPOX+Bev | 47  | 0.73 | (0.38 to 1.41) |                           |
| Yuki et al. 2020           |                   | 72  | 1.10 | (0.54 to 2.22) |                           |
| CM: CMS R package          |                   | 99  | 0.95 | (0.63 to 1.44) |                           |
|                            |                   | 90  | 0.72 | (0.46 to 1.13) |                           |
| FOLFOXIRI+Bev              | FOLFOX+Bev        | 308 | 0.60 | (0.38 to 0.95) |                           |
| Borelli et al. 2020        |                   |     |      |              |                           |
|                            | CM: NanoString    |     |      |              |                           |

### B

| Treatment                  | Control           | N   | HR   | 95% CI       | Overall survival          |
|----------------------------|-------------------|-----|------|--------------|---------------------------|
| Cap(±Mit)+Bev              | Cap               | 42  | 0.75 | (0.46 to 1.24) |                           |
| Mooi et al. 2018           |                   | 113 | 0.66 | (0.33 to 1.31) |                           |
| CM: CMS R package          |                   | 28  | 0.68 | (0.26 to 1.78) |                           |
|                            |                   | 54  | 2.12 | (1.04 to 4.34) |                           |
| FOLFIRI+Bev                | FOLFIRI           | 12  | 1.27 | (0.36 to 4.43) |                           |
| Del Rio et al. 2017        |                   | 18  | 0.50 | (0.11 to 2.38) |                           |
| CM: CMS R package          |                   | 14  | 3.81 | (0.40 to 36.7) |                           |
|                            |                   | 29  | 0.38 | (0.14 to 1.06) |                           |

### C

| Treatment                  | Control           | N   | HR   | 95% CI       | Overall survival          |
|----------------------------|-------------------|-----|------|--------------|---------------------------|
| FOLFIRI+Bev                | FOLFIRI+Cet       | 46  | 1.20 | (0.64 to 2.22) |                           |
| Stintzing et al. 2019      |                   | 130 | 1.16 | (0.79 to 1.72) |                           |
| CM: CMS R package          |                   | 36  | 1.75 | (0.81 to 3.70) |                           |
|                            |                   | 103 | 1.75 | (1.16 to 2.63) |                           |
| FOLFOX+Bev/                | FOLFOX+Bev/Cet    | 104 | 0.43 | (0.27 to 0.68) |                           |
| FOLFIRI+Bev                | FOLFIRI+Cet       | 242 | 1.61 | (1.16 to 2.22) |                           |
| Lenz et al. 2019           |                   | 68  | 0.92 | (0.52 to 1.61) |                           |
| CM: NanoString             |                   | 167 | 0.96 | (0.66 to 1.39) |                           |
| CAPOX+Bev                  | CAPOX+Bev+Cet     | 95  | 1.92 | (1.15 to 3.23) |                           |
| Trinh et al. 2017          |                   | 69  | 0.64 | (0.38 to 1.10) |                           |
| CM: Immunohistochemistry   |                   |     |      |              |                           |