INTRODUCTION

The relationship between depression and anxiety disorders has been the focus of attention for many mental health clinicians and researchers. Interest in this issue stems primarily from the high degree of comorbidity and correlation between these two conditions. Epidemiological studies have reported that the comorbidity rates between depression and anxiety disorders are from 30% to 75% (1, 2), and the correlation measured by self-reports has been placed between 0.50 and 0.80. The findings of research conducted in the 1970s and 1980s can be thought of as fitting into three categories: a) anxiety and depression differ quantitatively (i.e., unitary model); b) anxiety and depression differ qualitatively (i.e., dual model); c) combined anxiety and depression differ both quantitatively and qualitatively from either pure anxiety or pure depression (3). However, researchers have gradually begun to realize that the dispute over these three concepts is an unnecessary and unproductive one (4). In other words, it has been realized that it will be more effective and productive to determine exactly what the similarities and differences of anxiety and depression are, rather than debating whether the two conditions are quantitatively or qualitatively different. On the basis of this background, Clark and Watson (5) suggested a new concept, now commonly referred to as the tripartite model of anxiety and depression.

Clark and Watson hypothesized three relatively independent factors, positive affect (PA), negative affect (NA), and physiological hyperarousal (PH). They proposed that anxiety and depression are both characterized by the nonspecific distress of NA, whereas anxiety and depression each feature their own unique factors, high PH (anxiety) and low PA (depression). The tripartite model has rapidly received extensive empirical support (6, 7). This model, which was initially developed from adult empirical data, has been applied to both child and adolescent samples, and, in the past 15 yr, ever-increasing efforts have focused on the validity of the tripartite model of anxiety and depression in these younger age groups (8, 9).

For example, Joiner et al. (10) selected NA, PA, and PH construct items from the Children's Depression Inventory (CDI) and the Revised Children's Manifest Anxiety Scale (RCMAS), as well as the Children's Depression Inventory (CDI). The results indicated that the model fit of a three-factor model was superior to one- and two-factor models. In addition, the findings of discriminant analysis demonstrated that the correct classification rate with three factors of the tripartite model was superior to the classification rate achievable using CDI and RCMAS. In a departure from Clark and Watson's hypothesis, however, the correlations of three factors were significantly higher than had been expected. The results are discussed on the basis of cultural background.

Key Words: Child, Adolescent, Comorbidity, Anxiety, Depression

Validation Study of Tripartite Model of Anxiety and Depression in Children and Adolescents: Clinical Sample in Korea

Although the currently available literature has provided some empirical support for a tripartite model of child and adolescent anxiety and depression, one of the limitations of these studies was that they have been conducted in America, primarily with Caucasians. In order to make this model more applicable to diverse ethnic and cultural groups, this study used a tripartite model for child and adolescent anxiety and depression in Korea, using confirmatory factor analysis with logically selected items from the Revised Children's Manifest Anxiety Scale (RCMAS), as well as the Children's Depression Inventory (CDI). The results indicated that the model fit of a three-factor model was superior to one- and two-factor models. In addition, the findings of discriminant analysis demonstrated that the correct classification rate with three factors of the tripartite model was superior to the classification rate achievable using CDI and RCMAS. In a departure from Clark and Watson's hypothesis, however, the correlations of three factors were significantly higher than had been expected. The results are discussed on the basis of cultural background.

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(12) and Affect and Arousal Scale (13) have been developed, thus continuously providing further empirical support for the model.

To date, however, tests of the tripartite model have been conducted principally with Caucasian youths in America. To the best of our knowledge, only two studies of this subject with community adolescents have been conducted outside of America and these only provided partial empirical support (14, 15). As racial and ethnic differences exist in the symptom presentations of psychiatric disorders (16), more cross-cultural validation studies for tripartite model of anxiety and depression are necessary to broaden its application. Therefore, the objective of the present study was to validate the tripartite model of emotion to clinical samples in Korea. In order to compare our results with those of previous tests, we used items which had been logically selected by Joiner et al. (10).

We also attempted to determine how accurately anxiety and depression could be differentiated using the selected items, on the basis of the tripartite model.

**MATERIALS AND METHODS**

**Subjects**

The participants in this study included 150 youngsters (94 boys; 56 girls) with anxiety disorders and/or depression, all of whom were inpatients or outpatients in the Department of Psychiatry at a general hospital in Seoul, Korea. The mean age of the sample was 13.33 yr (SD=2.60; range=8 to 18). A psychiatrist had rendered their diagnoses after semi-structured clinical interviews with patients and their parents, on the basis of the DSM-IV (17). Information regarding the principal diagnoses of the samples is shown in Table 1. Twenty four patients exhibited other comorbid psychiatric disorders (12 individuals with attention deficit hyperactivity disorder, 3 with oppositional defiant disorder, 3 with conduct disorder, 2 pervasive developmental disorders, and 1 each of transient tic disorder, delusional disorder, intermittent explosive disorder, and developmental coordination disorder).

**Procedure**

On the initial visit, a psychiatrist interviewed with the patients and their parents. On the next visit, the patients individually completed the measurements used in this study. All of these assessments were conducted in an examination room within the hospital.

**Measures**

All participants completed children's depression scale (18, 19) and revised children's manifest anxiety scale (20, 21) which are most widely used self reports to measure depressive and anxiety symptoms. NA, PA and PH items were selected and analyzed for validation of tripartite model as were initially selected by Joiner et al. (10), in tests of the tripartite model conducted in previous studies. The selected NA, PA and PH constructs from CDI and RCMAS are shown in Table 2.

**Table 1. Principal diagnoses of sample**

| Diagnosis                          | Number |
|-----------------------------------|--------|
| Depressive Disorders              | 113    |
| Major depressive disorder         | 15     |
| Dysthymic disorder                | 23 (comorbid with 1 GAD) |
| Depressive disorder, NOS           | 75 (comorbid with 4 Anxiety disorder, NOS & 1 OCD) |
| Anxiety Disorders                  | 35     |
| Panic disorder                     | 2      |
| Social phobia                      | 1      |
| Obsessive-compulsive disorder      | 5      |
| Posttraumatic stress disorder      | 1      |
| Acute stress disorder              | 1      |
| Generalized anxiety disorder       | 1      |
| Anxiety disorder, NOS              | 24     |
| Mixed Anxiety-Depressive Disorder  | 2      |
| Total                              | 150    |

GAD, Generalized Anxiety Disorder; NOS, Not Otherwise Specified; OCD, Obsessive-Compulsive Disorder.

**Table 2. Selected NA, PA, and PH items from CDI and RCMAS by Joiner et al.**

| Measures | NA | (Low) PA | PH |
|----------|----|----------|----|
| CDI      | Sad all the time | Nothing is fun at all |
|          | Feel like crying every day | Do not want to be with people |
|          | Things bother me | Never have fun at school |
| RCMAS    | Worry a lot of the time | Others happier than I am |
|          | Afraid of a lot of things | Trouble getting breath |
|          | Get mad easily | Feel sick in stomach |
|          | Worry what parents say | Hands feel sweaty |
|          | Worry what others think |
|          | Worry when go to bed |
|          | Nervous |
|          | Worry about something bad |

NA, Negative Affect; PA, Positive Affect; PH, Physiological Hyperarousal; CDI, Children’s Depression Inventory; RCMAS, Revised Children’s Manifest Anxiety Scale.
Statistical analysis

In order to investigate the differences in CDI, RCMAS, and three constructs of tripartite model between anxiety and depression groups, t-test was conducted. For comparison of the fit of a tripartite model, confirmatory factor analyses were employed. The goodness of fit of the models was evaluated using the Comparative Fit Index (CFI), Tucker-Lewis Index (TLI), and the Root Mean Square Error of Approximation (RMSEA). The sequential chi-square difference test was used for the comparison of models. Additionally, a discriminant analysis was conducted in order to determine whether or not three factors (NA, PA, and PH) could be used to discriminate anxiety from depressive disorders. A \( p \) value of 0.05 or less was regarded as statistically significant. T-test and discriminant analysis were performed using SPSS 12.0 for Windows (SPSS, Chicago, IL, U.S.A.). Confirmatory factor analysis was computed with AMOS 4.0 (22), using maximum likelihood estimation method.

RESULTS

Group comparison

In the pure anxiety and depression comparisons, comorbid participants were excluded for group comparisons. Group differences were assessed via the comparison of means on the CDI, RCMAS and the subscales of the tripartite model.

|                  | Anxiety Disorders Only (n=35) | Depressive Disorders Only (n=107) | t (df=140) | \( p \) |
|------------------|------------------------------|----------------------------------|------------|------|
| CDI              | 15.33 (7.30)                 | 20.30 (7.94)                     | -3.28      | 0.001|
| RCMAS            | 15.16 (5.96)                 | 15.22 (6.33)                     | -0.05      | 0.963|
| NA (Low) PA      | 6.41 (3.31)                  | 5.94 (3.13)                      | 0.74       | 0.462|
| PH               | 1.89 (1.28)                  | 3.10 (1.74)                      | -3.82      | 0.000|
| PA               | 1.02 (1.05)                  | 1.49 (0.89)                      | -2.38      | 0.019|

Table 3. Group differences on CDI, RCMAS and subscales of tripartite models

| Model                  | \( \chi^2 \) | df | CFI | TLI | RMSEA |
|------------------------|-------------|----|-----|-----|-------|
| One-factor             | 207.81      | 135| 0.945 | 0.930 | 0.086 |
| Two-factor             | 209.94      | 134| 0.968 | 0.960 | 0.062 |
| NA-PA : PH             | 209.00      | 134| 0.969 | 0.960 | 0.061 |
| PA-PH : NA             | 196.06      | 134| 0.955 | 0.943 | 0.078 |
| NA-PH : PA             | 171.52      | 132| 0.984 | 0.979 | 0.045 |

Table 4. Goodness-of-fit indices for confirmatory factor analyses of five alternative models of the NA, PA, and PH

CONFIRMATORY FACTOR ANALYSIS

Confirmatory factor analyses were employed in the comparison of the fit of a tripartite model with (a) a unitary model and (b) three two-factor models, in which each two-constructs were collapsed into one factor. The results obtained using items selected on the basis of tripartite model showed that the model fit of the three-factor model was adequate, and was, in fact, significantly better than that of the one- and two-factor models: \( \chi^2(2, N=150)=24.54, p<0.01 \) (three-factor vs. one-factor), \( \chi^2(3, N=150)=36.29, p<0.01 \) (three-factor vs. two-factor; see Table 4). The correlations of three latent variables in the tripartite model were higher than expectation (NA:PH \( r=0.64 \); NA:PA \( r=0.46 \); PA:PH \( r=0.21 \)).

Discriminant analysis

We also conducted a discriminant analysis, in order to determine whether or not three factors (NA, PA, and PH) could be used to discriminate anxiety from depressive disorders. 142 patients were included in this analysis, and 8 subjects with comorbid anxiety and depression were excluded from participation. For comparison, we also conducted a discriminant analysis, using all items from the CDI and the RCMAS. Standardized function coefficients and correlation coefficients are shown in Table 5.

Table 5. Discriminant analysis for diagnosis of anxiety and depressive disorder groups
sification results showed that 80.0% of the anxiety disorders had been classified correctly, whereas 67.3% of the depressive disorders had been classified correctly. In the overall sample, 70.4% of the disorders were classified correctly. However, with the whole items of the CDI and RCMAS, the correct overall classification rate was 65.4%, and the rate for anxiety and depressive disorders was 68.6% and 65.4% each. In general, the correct classification rate of anxiety from depression was better with the NA, PA, and PH constructs than with all of the items of the CDI and RCMAS.

**DISCUSSION**

The result of this study, which was consistent with the findings of previous studies, reported that the model fit index of the three-factor structure was significantly superior to that of all of the other models. In addition, using the items logically selected from the CDI and RCMAS on the basis of tripartite model, we found there to be no significant differences between the anxiety and depression groups with regard to NA. However, the group with anxiety disorders showed more physiological hyperarousal than the group with depression. On the contrary, the group with depression showed lower in positive affect than the group with anxiety disorders. In addition, the findings of discriminant analysis indicated that the correct rate at which anxiety disorders could be differentiated from depression using the NA, PA, and PH constructs was higher than when using the whole scales, CDI and RCMAS. These findings provide further evidence supporting the validity of the tripartite model of emotion in a clinical sample of patients suffering from anxiety and depressive disorders.

To date, the validity of the tripartite model in children and adolescents has been verified by the results of prior empirical studies for a decade (8). However, almost all studies on this subject have been done in America, primarily with Caucasian youths. The results of this study, however, serve to broaden the application of the tripartite model to diverse ethnic and cultural groups.

Although the overall results of this study tend to support the tripartite model, three factors in this sample were highly correlated with one another, in contrast with Clark and Watson’s hypothetical explanation (5) that the tripartite factors are orthogonal dimensions.

Several explanations are possible for these unexpected results. The first explanation involves the sample used in this study. The subject group in this study was comprised of children and adolescents, all of whom had anxiety and depressive disorders which were sufficiently severe enough to warrant a visit to the department of psychiatry in a general hospital. Comorbidity with anxiety and depression is a very common phenomenon. Thus, it is not surprising that patients with either anxiety or depression also have increased tendencies toward the other symptoms, although these tendencies are not normally sufficiently severe to constitute additional diagnoses.

A second explanation concerns the scales used in this study, CDI and RCMAS. It has been repeatedly proposed that both the CDI and the RCMAS possess low discriminant validity, and are only used as measurements of general distress and NA, due to the fact that they were developed in the absence of a theoretical background (23, 24). The NA, PA, and PH items used in this study were selected from these scales, and thus there might be a high degree of overlap among the tripartite constructs, even though the items had been theoretically selected. By way of contrast, other investigations employing PANAS-C and AFARS, which were specifically designed to assess the tripartite model, indicated only a few correlations among factors, which is consistent with Clark and Watson’s hypothesis (5).

The third explanation is related to cultural differences. It is suggested that patients in non-Western cultures report somatic symptoms and deny psychological symptoms more frequently than patients in Western countries (25). In Korea, based on the traditional Confucian background of the culture, the direct expression of emotions is fairly restricted. In other words, the inhibition of emotional expression is a Korean cultural characteristic. Therefore, Koreans exhibit a tendency to express their negative feelings as somatic complaints. Considering Korean culture, it might be inferred that somatic symptoms tend to be more strongly associated with negative emotions, including anxiety and depressive symptoms, than in any other cultures, particularly Western cultures. The Korean cultural context may explain why the correlations of three constructs are higher than was found in the prior studies, which were conducted primarily in America, and in Clark and Watson’s hypothesis; thus, cultural background should be considered to some degree when applying the tripartite model to other cultures. In order to resolve issues regarding the relationships between the three constructs in the tripartite model, three possibilities might be considered. Additional research with multiple scales, and in other countries, will be required for a solid solution.

To the best of our knowledge, this study is the first to investigate the tripartite model of anxiety and depression in Asia. The findings from this study serve to expand the applicability of the tripartite model. However, it should be noted that this study suffers from several inherent methodological limitations. Structured interviews for diagnosis were not conducted in this study, although semi-structured clinical interviews were conducted by a psychiatrist. Besides, due to the fact that we used only children and adolescents’ self-reports as an informant, we abrogated the benefits which would have been generated from the use of multiple informants. In addition, it should be noted that it was one of the limitations that the subject number of anxiety disorders was relatively small in this study. Considering the limitations of this study, though, more research with diverse cultures will be required in order to facilitate the application of the tripartite model in the future.
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