A Pharmacometric Framework for Axitinib Exposure, Efficacy, and Safety in Metastatic Renal Cell Carcinoma Patients

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The relationships between exposure, biomarkers (vascular endothelial growth factor (VEGF), soluble VEGF receptors (sVEGFR)-1, -2, -3, and soluble stem cell factor receptor (sKIT)), tumor sum of longest diameters (SLD), diastolic blood pressure (dBp), and overall survival (OS) were investigated in a modeling framework. The dataset included 64 metastatic renal cell carcinoma patients (mRCC) treated with oral axitinib. Biomarker timecourses were described by indirect response (IDR) models where axitinib inhibits sVEGFR-1, -2, and -3 production, and VEGF degradation. No effect was identified on sKIT. A tumor model using sVEGFR-3 dynamics as driver predicted SLD data well. An IDR model, with axitinib exposure stimulating the response, characterized dBp increase. In a time-to-event model the SLD timecourse predicted OS better than exposure, biomarker- or dBp-related metrics. This type of framework can be used to relate pharmacokinetics, efficacy, and safety to long-term clinical outcome in mRCC patients treated with VEGFR inhibitors. (ClinicalTrial.gov identifier NCT00569946.)

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Study Highlights

 WHAT IS THE CURRENT KNOWLEDGE ON THE TOPIC? ☑ A modeling framework in sunitinib-treated gastrointestinal stromal tumors identified circulating biomarkers and adverse effects as better predictors of overall survival (OS) than tumor size (SLD). Similar relationships may be of value for predicting OS in metastatic renal cell carcinoma (mRCC) patients treated with axitinib.

 WHAT QUESTION DID THIS STUDY ADDRESS? ☑ The relationships between axitinib exposure, biomarkers related to VEGFR inhibition, hypertension (the most common adverse effect for axitinib), SLD, and OS were investigated in axitinib-treated Japanese mRCC patients.

 WHAT THIS STUDY ADDS TO OUR KNOWLEDGE ☑ Early changes in soluble VEGFR-3 could forecast tumor response. This analysis is one of the first to demonstrate SLD dynamics as a predictor of OS, which was better than biomarker- or hypertension-related metrics or tumor size change at a specific week.

 HOW MIGHT THIS CHANGE DRUG DISCOVERY, DEVELOPMENT, AND/OR THERAPEUTICS? ☑ The modeling framework can be used as a template to leverage data collected during oncology clinical trials when developing new targeted therapies, facilitate identification of predictors for long-term clinical outcome, and select the most promising dosing schedules.

In metastatic renal cell carcinoma (mRCC) the vascular endothelial growth factor (VEGF) is typically overexpressed and mRCC is predominantly refractory to traditional cytotoxic chemotherapies. Several first-line treatment alternatives with targeted therapies exist, including the tyrosine kinase inhibitors (TKIs) sunitinib and pazopanib.1 However, patients often develop biological resistance and receive second-line treatment.2 Axitinib is a potent and selective oral TKI targeting the VEGF receptors (VEGFR) 1, 2, and 3 and primarily displays antiangiogenic activity. The drug is approved in Europe, the United States, Japan, and elsewhere for the treatment of advanced renal cell carcinoma (RCC) after failure of one prior systemic therapy,3 and is currently a preferred choice as second-line therapy for patients progressing after first-line therapy.3 Moreover, axitinib has shown clinical activity in first-line mRCC in recent phase II and III trials.4,5 Axitinib is approved at a starting dose of 5 mg twice daily (b.i.d.) and dose increase or reduction is recommended based on individual safety and tolerability, including increased blood pressure (BP). Dose titration enables patients with good tolerability at a 5 mg starting dose to reach higher exposures6 and results in a better objective response rate.6

The conventional Response Evaluation Criteria in Solid Tumors (RECIST), which are based on a categorization of the response seen on the sum of longest diameters (SLD), were designed to evaluate therapeutic efficacy of cytotoxic agents.7 However, RECIST may not reflect the clinical benefit of antiangiogenic drugs for which tumor shrinkage may be limited or delayed.8 Increases in blood pressure are common after initiation of anti-VEGF therapy9 and have been proposed as an independent predictor for overall survival (OS) and progression-free survival (PFS) in axitinib-treated mRCC patients10 and axitinib- and other TKI-treated solid tumors,11
including sunitinib-treated gastrointestinal stromal tumors (GIST).\textsuperscript{12,13} Optimal axitinib exposure, leading to best achievable long-term outcome, may, however, differ among mRCC patients and dose selection cannot likely be solely based on pharmacokinetics (PK) or BP measurements.\textsuperscript{6} Increases in VEGF and decreases in the soluble fragments of its receptors (sVEGFR-1, -2, and -3) have been suggested as biomarkers of angiogenesis inhibition and predictors for clinical response in RCC treated with TKIs,\textsuperscript{14} including axitinib.\textsuperscript{15,16} A better understanding of the relationships between axitinib exposure, plasma biomarkers, BP, SLD, and long-term clinical outcome can be valuable for identifying robust pharmacodynamic (PD) biomarkers and guide treatment decisions.

By integrating quantitative knowledge on anticancer drugs’ safety and efficacy, pharmacometric modeling has shown value in guiding oncology clinical trial design and rational dose selection, thereby optimizing benefit/risk management for cancer patients.\textsuperscript{17–20} As an example, in an overarching modeling framework Hansson et al. elucidated the relations between drug exposure, the timecourse of circulating biomarkers (VEGF, sVEGFR-2, and sVEGFR-3, and soluble stem cell factor receptor (sKIT)), SLD, adverse effects (fatigue, hand-foot syndrome, neutropenia, and hypertension), and OS in sunitinib-treated GIST patients.\textsuperscript{12,21} Increased VEGF, decreased sVEGFR-2, sVEGFR-3, and sKIT concentrations and diastolic BP (dBP) elevation were dependent on sunitinib exposure and dosing schedule. Sunitinib exposure together with sVEGFR-3 and sKIT dynamics predicted the SLD timecourse. A smaller baseline SLD and larger sVEGFR-3 decrease over time were associated with longer OS. Alternatively, hypertension and neutropenia could be used as predictors for OS.

In the present work, the relationships between axitinib exposure, the timecourses of potential biomarkers (VEGF, sVEGFR-1, -2, and -3, sKIT), SLD, dBP, and OS in axitinib-treated Japanese mRCC patients were explored and quantified using pharmacometric models.

### METHODS

#### Patients and data

During the development of axitinib, biomarker data were collected from 64 Japanese cytokine-refractory mRCC patients involved in a single-arm, open-label, multicenter phase II study.\textsuperscript{22} Axitinib was administered in 4-week cycles at a starting dose of 5 mg b.i.d. In eligible patients axitinib dose was increased by 2–3 mg b.i.d. up to 10 mg b.i.d. every 2 weeks or more ($n = 5$), or decreased to 2 mg b.i.d. ($n = 41$) based on tolerability (BP and other nonhematologic adverse effects) and dosing history was recorded. The major reason for dose reduction or treatment discontinuation/interruption was proteinuria (28%). SLD was measured according to RECIST 1.0. Biomarkers, SLD, dBP, and OS assessment times are summarized in Table 1. This study was conducted in accordance with the Declaration of Helsinki, the International Conference on Harmonisation guidelines on Good Clinical Practice, and applicable local regulatory requirements and laws. All participants provided informed consent. The study protocol was approved by an institutional review board at each site.

#### Model development

Nonlinear mixed effect models were developed using NONMEM software v. 7.3.\textsuperscript{23} Parameters were estimated using the first-order conditional estimation method with interaction (FOCEI), and for dropout and OS analysis, the Laplacian estimation method. R v. 3.1.1, the R-based package Xpose v. 4, Perl-speaks-NONMEM (PsN) toolkit v. 4, and Piraña v. 2.9.0 were used for data pre- and postprocessing, graphical visualization, and model diagnostics.\textsuperscript{24}

Model selection was based on goodness-of-fit plots and the objective function value (OFV, $-2 \log$-likelihood). A significance level of $P < 0.05$ as assessed by the OFV difference (dOFV) was used to discriminate between nested models. The predictive performance of the biomarkers, SLD, and dBP models was assessed using (prediction-corrected) visual predictive checks ((pc)VPCs),\textsuperscript{25} where 95% confidence intervals (CIs) derived from 500 simulated

### Table 1 Summary of study assessments and available data

| Variable | Per protocol assessment time (study day) | Available data (n; follow-up duration in days, median [range]) |
|----------|------------------------------------------|------------------------------------------------------------------|
| VEGF, sVEGFR-1, -2, -3, sKIT | Cycle 1: pre-dosing | $n = 436$ for each biomarker; $168 [32-624]^a$ |
| | Cycle 2-7: day 1 | |
| | EoT/discontinuation | |
| Sum of longest diameters | Cycle 1: pre-dosing | $n = 476; 337 [36-731]^a$ |
| | Subsequent odd no. cycles: day 1 | |
| | EoT/discontinuation | |
| Diastolic blood pressure | Cycle 1: pre-dosing, day 8, 15, 22 | $n = 308; 29 [21-29]^a$ |
| | Cycle 2-4: day 1, 15 | |
| | Cycle >4: day 1 | |
| | EoT/discontinuation | |
| Overall survival | Until EoT/discontinuation and every 6 months thereafter | $16/48$ deaths/censored; $457 [85-781]$ |

\textsuperscript{a}Summary statistics on follow-up duration exclude one patient with biomarker data available at baseline only.

\textsuperscript{b}Summary statistics on follow-up duration exclude two patients with tumor data available at baseline only.

\textsuperscript{c}Only data from the first month (all visits in Cycle 1 and day 1 in Cycle 2) were modeled.
datasets were compared to the observed data. Kaplan–Meier VPCs, comparing the 95% CI derived from 200 simulations to the observed time-to-event (TTE) data, were used to evaluate the dropout and OS model performance. Relative standard errors (RSE) of parameter estimates were obtained from the NONMEM Sandwich matrix for continuous data and from the R matrix for dropout and OS models.

Exponential and additive interindividual variability (IIV) were evaluated as appropriate. Semiparametric distributions were tested when indicated graphically.26 Residual unexplained variability (RUV) was evaluated for all continuous data models using additive, proportional, or combined error models.

Pharmacokinetics
Empirical Bayes estimates (EBEs) of apparent clearance ($CL/F$) obtained from a published population PK model10 were used to calculate the daily area under the concentration–time curve $AUC_{daily} = \frac{Dose_{daily}}{CL/F}$, where $Dose_{daily}$ is the daily dose accounting for dose increases and reductions. Since axitinib has a relatively short typical elimination half-life (3 h in Japanese patients),10 $AUC_{daily}$ was assumed to be 0 on days off-therapy.

Biomarker models
Indirect response (IDR) models27 where axitinib inhibits sVEGFR-1, -2, and -3 and sKIT production, and VEGF degradation, were investigated. Linear, maximal effect ($E_{max}$) and sigmoidal $E_{max}$ drug effects driven by $AUC_{daily}$ were evaluated. Linear disease progression functions were explored to describe potential changes in biomarkers in the absence of drug. Models for each biomarker were developed separately before being combined into a joint model to explore correlations between model parameters.

Tumor size model
Tumor models where axitinib induces a decrease in SLD were investigated. Zero-order and first-order tumor growth were evaluated. The axitinib effect on SLD was driven either directly by axitinib exposure ($AUC_{daily}$), or indirectly by individual model-predicted changes in the different biomarkers (absolute value, or absolute or relative change from baseline). An approach similar to population PK parameters and data (PPP&D) was adopted,28,29 i.e., population biomarker parameters were fixed while individual biomarker parameters were predicted simultaneously with SLD parameters based on both biomarker and SLD data. Drivers were tested alone and in combination. An exponential decay in drug effect describing potential tumor regrowth was investigated.

As dropout from tumor measurements may not be completely random, a logistic regression model was developed to mimic varying measurement durations in the SLD simulations. Investigated predictors in the dropout model included the time since start of treatment, $AUC_{daily}$, the observed baseline SLD and predicted SLD at the time of evaluation, and progressive disease (PD, yes/no) defined as a 20% increase in SLD from nadir. During simulations, dosing records were imputed based on the last observed dose until the time of last observed tumor assessment in the study.

Diastolic blood pressure model
According to protocol, new or additional antihypertensive treatments would be prescribed to patients having dBP elevation during axitinib therapy. Therefore, only data from the first treatment cycle were modeled. IDR models were investigated with axitinib stimulating the production of dBP response through linear, power, or (sigmoidal) $E_{max}$ drug effects driven by $Dose_{daily}$ or $AUC_{daily}$.

Overall survival model
Parametric TTE models were developed for OS data. Exponential, Weibull, Gompertz, log-normal, and log-logistic distributions were investigated to describe the baseline hazard. Predictors were tested one by one and thereafter in combination. Individual parameters were used to compute the model-based predictors. Due to model instability their uncertainty could not be accounted for.30 Evaluated baseline predictors included Eastern Cooperative Oncology Group (ECOG) performance status, demographics (age, sex, body weight), and model-predicted baseline biomarkers, dBP and SLD values. Time-varying predictors included $Dose_{daily}$, $AUC_{daily}$, the predicted timecourse, absolute and relative change from baseline in biomarkers and SLD, and the derivative of SLD predicted timecourse. Dose and time-varying predictors were extrapolated based on the last recorded dose assuming that patients were treated with axitinib until death or censoring, as the protocol supported treatment continuation in case of clinical benefit (no new lesion and SLD smaller than at baseline) despite progressive disease according to RECIST. Additionally, model-predicted relative changes from baseline in biomarkers at week 4, in dBP at weeks 2 and 4, and in SLD at week 8, the maximum absolute dBP during cycle 1 and a dBP greater than 90 mmHg during cycle 1 (yes/no) were evaluated.

Censoring, defined as loss to follow-up or nonoccurrence of death at the end of the study, was described by a competing hazard function to account for varying follow-up durations.

RESULTS
A schematic representation of the final modeling framework is depicted in Figure 1.

Patients and data
Patients were treated with axitinib for a median of 51 weeks (range, 1.7–104). Available biomarker, SLD, dBP, and OS data are summarized in Table 1. Four sVEGFR-1 and four sVEGFR-3 concentration values (<1%) were below the limit of quantification and omitted from the analysis dataset. Two patients had SLD data available at baseline only. At the end of the follow-up period, 16 patients had died and 48 patients were censored from OS analysis.

Biomarkers’ models
Log-transformed biomarker data were well described by IDR models where axitinib inhibits VEGF degradation (Eq. 1) and sVEGFR-1, -2, and -3 production (Eq. 2) (Figure 1). A linear

\[ AUC_{daily} = \frac{Dose_{daily}}{CL/F} \]

\[ TTE_{model} = \frac{1}{\lambda(TTE)} \]

\[ \lambda(TTE) = \lambda_0 \exp \left( \beta_1 \right) \]

\[ \lambda_0 = \frac{1}{\text{OS Median}} \]

\[ \beta_1 = \text{baseline hazard parameter} \]

\[ \text{OS Median} = 26.0 \text{ weeks} \]

\[ \text{OS Hazard Ratio} = 0.8 \]
time-dependent disease progression component with slope $\alpha$ described a drug-independent VEGF increase, whereas no disease progression was identified for sVEGFR-1, -2, and -3.

Drug effects were described by inhibitory $E_{\text{max}}$ (VEGF, sVEGFR-1 and -3) or sigmoidal $E_{\text{max}}$ models (sVEGFR-2) assuming that maximum inhibition can be achieved ($I_{\text{max}} = 1$).

$$\frac{dA}{dt} = R_{\text{in}} \cdot (1 + \alpha \cdot t) - k_{\text{out}} \cdot \left(1 - \frac{I_{\text{max}} \cdot AUC_{\text{daily}}}{AUC_{50} + AUC_{\text{daily}}} \right) \cdot A(t)$$

(1)

$$\frac{dA}{dt} = R_{\text{in}} \cdot \left(1 - \frac{I_{\text{max}} \cdot AUC_{\text{daily}}}{AUC_{50} + AUC_{\text{daily}}} \right) - k_{\text{out}} \cdot A(t)$$

(2)

$k_{\text{out}}$ is the first-order rate constant for the biomarker degradation, expressed as $k_{\text{out}} = 1/MRT$, where MRT is the biomarker mean residence time in plasma. $R_{\text{in}}$ is the zero-order rate constant for biomarker production, calculated as $R_{\text{in}} = k_{\text{out}} \cdot \text{Base}$ with Base being the baseline biomarker concentration. $\gamma$ is the Hill coefficient and $AUC_{50}$ the $AUC_{\text{daily}}$ leading to half $I_{\text{max}}$. No axitinib drug effect was identified on sKIT and the data were best described by a linear and constant change over time (Supplementary Material). The sKIT model was therefore not included in the joint biomarker model, nor were sKIT-related predictors tested on SLD or OS.

When VEGF, sVEGFR-1, -2, and -3 were modeled jointly, large correlations (80–99%) were identified between individual $AUC_{50}$ for sVEGFR-1, -2, and -3 and the IIV magnitudes were similar; hence, a common IIV term was used. Moreover, the $AUC_{50}$ could be shared for sVEGFR-2 and -3 without worsening the model fit. Parameter estimates and their uncertainty are reported in Table 2. $AUC_{50}$ values were in the range of observed $AUC_{\text{daily}}$ (31.95–1,861 µg/h/L), with VEGF being most sensitive to axitinib ($AUC_{50}$ of 354 vs. 717–1,380 µg/h/L for the other biomarkers). VEGF and sVEGFR-1 typically displayed fast turnover (MRT of 0.722 and 0.624 days, respectively) compared to sVEGFR-2 and -3 (MRT of 19.7 and 5.76 days, respectively). A common additive RUV term for all four biomarkers was applied to account for that the biomarkers were sampled at the same time. All model parameters were estimated with reasonable uncertainty (<34% RSE), except for $MRT$ of sVEGFR-1 (69% RSE), for which the 95% CI obtained from sampling importance resampling was 0.0444–1.58 days.
Table 2 Parameter estimates and their uncertainty for the final joint biomarker model

| Biomarker | VEGF | sVEGFR-1 | sVEGFR-2 | sVEGFR-3 |
|-----------|------|----------|----------|----------|
| Base (pg/mL) | 65.0 (7.8) | 83.5 (2.9) | 8,850 (2.8) | 19,500 (6.5) |
| MRT (days) | 0.722 (25) | 0.624 (69) | 19.7 (17) | 5.76 (12) |
| AUC_{tot} (μg h/L) | 354 (13) | 1,380 (13) | 717 (8.6) | 717 (8.6) |
| γ | 1 FIX | 1 FIX | 1 FIX | 1 FIX |
| x (year^{-1}) | 0.650 (28) | 0.193 (5.3) | 0.162 (14) | 0.263 (6.5) |
| Common RUVd | 0.0593 (26)* | 0.0593 (26)* | 0.0593 (26)* | 0.0593 (26)* |

VEGF, vascular endothelial growth factor; sVEGFR-1, 2, 3, soluble vascular endothelial growth factor receptor 1, 2, 3; RSE, relative standard error; IIV, inter-individual variability; CV, coefficient of variation; Base, baseline biomarker concentration; MRT, mean residence time; Imax, maximal inhibitory effect; AUC_{tot}, axitinib area under the concentration-time curve giving half of the maximal effect; γ, Hill coefficient; x, slope of the disease progression; RUV, residual unexplained variability.

*The 95% confidence interval obtained from sampling importance resampling was 0.0444–1.58 day.

The IIV in AUC_{tot} for VEGFR-1, 2, and 3 was quantified using a common variability term.

Common AUC_{tot} parameter for sVEGFR-2 and 3.

Expressed as standard deviation on log-scale.

Common RUV for all four biomarkers.

Figure 2 Prediction-corrected visual predictive checks of the final biomarker models based on 500 simulations. Median (solid line), 5th, and 95th percentiles (dashed lines) of the observed data (solid circles) are compared to the 95% confidence intervals (shaded areas) for the median, 5th, and 95th percentiles of the simulated data. VEGF, vascular endothelial growth factor; sVEGFR-1, -2, -3, soluble VEGF receptor 1, 2, 3.
Tumor size model
A tumor size model with an underlying first-order growth process with rate constant $K_G$ best described SLD data (Eq. 3, Figure 1). The individual model-predicted relative change from baseline over time in sVEGFR-1 ($sVEGFR1_{rel}(t)$, OFV = $-339.99$) and sVEGFR-3 ($sVEGFR3_{rel}(t)$, OFV = $-339.01$) were better drivers of SLD response than all other investigated predictors, including $AUC_{daily}$ (OFV = $-327.06$) and $sVEGFR-2$ ($sVEGFR2_{rel}(t)$, OFV = $-326.38$). Despite $sVEGFR-1_{rel}(t)$ had a one unit lower OFV, $sVEGFR-3_{rel}(t)$ was chosen to drive the SLD response in the final model, given the large uncertainty in MRT in the sVEGFR-1 model and for consistency with the published modeling framework in sunitinib-treated GIST. When $sVEGFR3_{rel}(t)$ was included in the model, none of the other predictors further improved the model fit.

$$\frac{dSLD}{dt} = K_G \cdot SLD(t) - k_{sVEGFR3} \cdot sVEGFR3_{rel}(t) \cdot e^{-\lambda t} \cdot SLD(t) \quad (3)$$

$k_{sVEGFR3}$ is the tumor size reduction rate constant related to sVEGFR-3 response and $\lambda$ the tumor resistance/regrowth appearance rate constant. The observed baseline tumor size (SLD(0)) was included as a covariate (i.e., not as a dependent variable), acknowledging the same RUV as for postbaseline observations (B2 method). The SLD data contained little information on $K_G$, and therefore its value and uncertainty (3.74 $10^{-3}$ week$^{-1}$, 6.64% RSE) obtained from a simplified tumor growth inhibition model developed using data from several clinical studies in RCC were used as informative prior for $K_G$ using the NONMEM SPRIOR subroutine.

The probability of dropping out was estimated to increase with the occurrence of PD ($\lambda_{PD} = 1.22$), higher SLD at the time of evaluation ($\lambda_{SLD} = 0.00282$ mm)$^{-1}$, increasing time since start of study ($\lambda_{Time} = 0.00371$ day)$^{-1}$ and decreasing $\text{AUC}_{daily}$ ($\lambda_{AUC} = -0.00529$ L$^{-1}$ h$^{-1}$$^{-1}$) (Equations in Supplementary Material). The final SLD and dropout model parameters and their uncertainty are reported in Table 3. The VPCs of the final SLD model accounting for dropout (Figure 3) demonstrate a good predictive performance of the model.

Diastolic BP model
An IDR model with a stochastic effect of axitinib on response production ($R_{\text{in},\text{dBp}}$) with an $AUC_{\text{daily}}$-driven $E_{\text{max}}$ model, parameterized as a maximal effect $E_{\text{max dBp}}$ and a slope parameter ($S_{0,\text{dBp}} = E_{\text{max dBp}} / AUC_{\text{50 dBp}}$) best described dBp data (Eq. 4, Table 3, Figure 1). $E_{\text{max dBp}}$ was estimated to 0.197 and $S_{0,\text{dBp}}$ to 0.00127 L$^{-1}$ h$^{-1}$$^{-1}$, corresponding to an $AUC_{\text{50 dBp}}$ of 155 µg/hL. A Box-Cox transformation with an estimated shape parameter of −5.42 was applied to baseline dBp ($\text{dBp}_{0}$) IIV to account for the skewed random effects distribution.

$$\frac{dBp}{dt} = R_{\text{in},\text{dBp}} \cdot \left(1 + \frac{E_{\text{max dBp}} \cdot S_{0,\text{dBp}} \cdot AUC_{\text{daily}}}{E_{\text{max dBp}} + S_{0,\text{dBp}} \cdot AUC_{\text{daily}}} - k_{\text{out},\text{dBp}} \cdot dBp(t) \right) \quad (4)$$

The rate constant for the loss of response is defined as $k_{\text{out},\text{dBp}} = 1 / MRT_{\text{dBp}}$, where $MRT_{\text{dBp}}$ is the mean turnover time (4.92 days) associated with dBp response, and $R_{\text{in},\text{dBp}} = k_{\text{out},\text{dBp}} \cdot dBp_{0}$. $AUC_{\text{daily}}$-predicted dBp response better than $Dose_{\text{daily}}$ (DOFV = 7.86). No IIV was identified on drug effect parameters ($E_{\text{max dBp}}$ and $S_{0,\text{dBp}}$). An additive model best described the RUV. The pcVPC (Figure 3) shows that the model well predicted the overall increase in dBp and the variability during the first month of treatment.

Overall survival model
OS baseline hazard was best described by a log-logistic distribution with a shape parameter $\gamma$ and scale parameter $\beta_{\text{SLD}}$. In the univariate analysis, the best predictor was SLD timecourse (SLD(t), DOFV = −26.9), followed by SLD baseline (DOFV = −13.3). Absolute change in SLD over time, the derivative of SLD predicted timecourse, SLD...
relative change from baseline at week 8 (i.e., corresponding to tumor size ratio week 8), biomarker-related predictors (VEGFR-2(t), VEGF(t), absolute change in sVEGFR-1(t) and relative change in VEGF from baseline at week 4) and dBP relative change at week 2 also resulted in statistically significant OFV drops; however, these drops were all driven by single individuals. Using both baseline SLD and absolute change in SLD(t) resulted in an OFV drop similar to SLD(t) (dOFV = −27.5) but required one extra parameter. When SLD(t) was included in the model, none of the other predictors further improved model fit. The final OS model is described by Eqs. 5 and 6, and parameter estimates and their uncertainty are reported in Table 3.

\[
h(t) = \frac{\psi^2 \cdot t^{-1}}{\gamma \cdot (1 + (\psi \cdot t)^2)} \cdot e^{\beta_{\text{SLD}} \cdot \text{SLD}(t)} \\
\psi = e^{-\beta_0}
\]  

\(\gamma\) was estimated to 0.298, meaning that the hazard, in the absence of changes in SLD(t), initially rises with time before decreasing monotonically (\(\gamma < 1\)). \(\beta_{\text{SLD}}\) is the coefficient for SLD(t) effect on the hazard, estimated to 0.0115 reflecting a 12% increase in hazard for a 10 mm increase in SLD. A competing log-logistic function described the hazard of being censored. Kaplan–Meier VPCs for OS (Figure 4) and censoring (Supplementary Material) show adequate predictive properties of the OS model.

**DISCUSSION**

In this pharmacometric framework (Figure 1) we investigated the relationships between drug exposure, soluble biomarkers, tumor size, hypertension, and OS following axitinib treatment in cytokine-refractory mRCC. Model-predicted sVEGFR-3 dynamics predicted tumor size better than axitinib exposure. The SLD timecourse was the best predictor of OS, with an estimated hazard ratio (HR) of 1.12 (95% CI of 1.08–1.17) for every 10 mm SLD increase.

The biomarker timecourses were successfully characterized by IDR models where axitinib inhibited VEGF
degradation and sVEGFR-1, -2, and -3 production. These model structures are consistent with previously published models in sunitinib-treated healthy volunteers, and metastatic colon cancer (mCC). The mechanisms behind these biomarker modulations following VEGFR inhibitor administration have not been fully elucidated. The hypothesis that the VEGF increase observed following VEGFR-2 inhibition may arise from a reduction in VEGF blood clearance is supported by our model. sVEGFR-2 decrease may result from a ligand-induced downregulation of VEGFR-2 from the cell surface, as shown in vitro. No axitinib effect was identified on sKIT, confirming previous findings that axitinib has negligible effect on the stem cell factor receptor KIT and acts as a selective VEGFR inhibitor. These results differ from those of the study in GIST patients (mostly Caucasian) treated with sunitinib, which markedly increases sKIT levels and is less specific to VEGFRs. Results from the joint biomarker model show that in this patient population, axitinib more potently inhibited sVEGFR-2 and -3 (AUC50 of 717 ± μg/L) than sVEGFR-1 (AUC50 of 1,380 ± μg/L). Consistent with the in vitro findings, sVEGFR-2 and -3 AUC50 values were lower than for sunitinib in GIST. The estimated typical VEGF and sVEGFR-2 baseline values were similar to those in healthy volunteers, GIST, and mCC, whereas sVEGFR-3 baseline was about 3-fold lower than in GIST (63,900 pg/mL) but similar to mCC (21,900 pg/mL). The typical sVEGFR-3 MRT was shorter in mRCC (5.76 days) than in GIST (16.7 days), denoting a faster turnover rate in mRCC but resulting in similar production rate constant (Rm).

The SLD reduction, seen in most patients after the start of therapy, was described by a tumor model allowing for axitinib-induced tumor shrinkage. Tumor resistance was estimated to develop at a faster rate in axitinib-treated mRCC than in sunitinib-treated GIST (0.101 vs. 0.0217 week⁻¹). The larger the sVEGFR-3 decrease, the more profound was the predicted tumor shrinkage. In sunitinib-treated GIST, sVEGFR-3 and sKIT reduction, as well as larger sunitinib AUC, were predictive of SLD decreases. In sunitinib-treated HCC, sVEGFR-2 absolute change from baseline was used as a driver for SLD response in a tumor growth inhibition model. It should be noted that sVEGFR-3 was not evaluated in that study. Our results add evidence to the fact that changes in PD biomarkers, which can easily be measured in plasma, can help to better understand and forecast tumor response in several cancer types treated with VEGFR inhibitors.

BP increases in most patients treated with angiogenesis inhibitors targeting the VEGF pathway, but the underlying pathophysiological mechanisms are not fully understood. Proposed mechanisms include reduction in nitric oxide production, increased prohypertensive agents expression, renin-angiotensin system activation, microvascular rarefaction, oxidative stress, pressure-natriuresis system, and arterial stiffness. Whereas a BP increase is generally promptly managed by antihypertensive therapies or TKI dose reduction, early dBP elevation after treatment initiation could be an easy-to-measure biomarker reflecting effective VEGF inhibition. The empirical IDR dBP model presented here identified a nonlinear exposure–response relationship. Chen et al. characterized ambulatory dBP data in mRCC patients monitored over a 24-h period predosing and 4 and 15 days after axitinib treatment initiation, using an IDR model where two cosine functions described dBP diurnal changes. In our dataset, the dBP measurements were performed weekly and diurnal variations could not be accounted for. The Eₘₐₓ estimate in the present analysis was similar to their findings (19.7% vs. 20.8%). The turn-over rate (Kₒᵤₐₛ = 1/MRT) differed, however, with an estimate of 0.203 day⁻¹ in our model vs. 0.254 h⁻¹ in the ambulatory setting. The previous analysis implies a time to dBP steady-state of 14 h that would not explain the current data, where dBP reaches steady-state after 1–2 weeks. This discrepancy may be explained by differences in study design (e.g., observation frequency).

The pharmacometric framework developed in sunitinib-treated GIST identified baseline SLD and sVEGFR-3 dynamics as the best predictors for OS, while here the SLD timecourse was the best predictor. These differences may be due to a discrepancy in tumor dynamics in the two patient populations: more axitinib-treated mRCC patients achieve complete or partial tumor response compared to sunitinib-treated GIST, for which stable disease is more frequent. An alternative model in GIST, absolute neutrophil count (ANC) reduction combined with dBP increase and baseline SLD were predictive of longer OS. No strong association between dBP and OS was identified here, which contrasts with previous findings for axitinib-treated mRCC, where a higher maximum dBP during the first 8 weeks of treatment was related to longer OS. The lack of information on antihypertensive therapy in later cycles prohibited a more thorough analysis of dBP-OS relationships. However, since steady-state was reached within the first cycle, later changes are primarily expected to be related to changes in dose. Since axitinib rarely induces neutropenia, ANC was not included in our analysis.

In the statistical analysis of long-term OS data in axitinib-treated Japanese mRCC patients, a baseline ECOG of 0 and a greater sVEGFR-2 reduction were associated with longer OS. In our analysis, ECOG or sVEGFR-2 could not predict OS data, which may be explained by the shorter follow-up period for OS (maximum 112 vs. ~285 weeks) in our analysis to avoid confounding effects of subsequent therapies after axitinib discontinuation. In a population analysis in first-line or refractory RCC patients treated with temsirolimus, sunitinib, or axitinib, time-to-tumor-growth (TTG) could predict OS. However, TTG can suffer from time-dependent bias. Moreover, a large variety of tumor profiles may lead to the same TTG, and TTG ignores the extent of tumor shrinkage. For these reasons, TTG was not tested on OS here. Instead, the tumor timecourse was identified to be the best predictor for OS, as previously suggested.

A potential limitation of our analysis is that it exclusively included data from Japanese patients; validation in a non-Japanese population may be required. Although no PK differences are expected between Japanese and non-Japanese patients, ethnic/racial differences in axitinib efficacy and safety may exist.
In summary, the sVEGFR-3 relative decrease over time was identified as a driver of tumor dynamics, which in turn was predictive of OS in axitinib-treated mRCC patients. Together with previous findings in sunitinib-treated GIST, our results support the use of sVEGFR data to better anticipate tumor response in patients treated with VEGF pathway inhibitors. In contrast to sunitinib-treated GIST, BP and biomarkers dynamics were not as good predictors of OS as SLD timecourse. Using the tumor timecourse is indeed more attractive from a theoretical standpoint than summary variables, such as tumor size ratio at a specific day or TTG. This type of overarching pharmacometric framework allows for leveraging clinical trial data and improved understanding of the relationships between drug exposure, potential plasma biomarkers, tumor size, frequently observed adverse effects, and long-term outcome, and can serve as platforms for identifying safe and efficacious dosing regimens through simulations.

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Conflict of Interest. M.A.A. is an employee of Pfizer Ltd. M.O.K. and L.E.F. have acted as paid consultants to Pfizer Ltd. As Deputy Editor-in-Chief of CPT: Pharmacometrics & Systems Pharmacology, L.E.F. was not involved in the review or decision process for this article. E.S. declared no conflict of interest.

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