Naive mean field approximation for image restoration

Hayaru Shouno
Graduate School of Human Culture, Nara Women’s University

Koji Wada
Graduate School of Science and Engineering, Saitama University

Masato Okada
Brain Research Institute, RIKEN
(Dated: October 27, 2018)

We attempt image restoration in the framework of the Bayesian inference. Recently, it has been shown that under a certain criterion the MAP (Maximum A Posteriori) estimate, which corresponds to the minimization of energy, can be outperformed by the MPM (Maximizer of the Posterior Marginals) estimate, which is equivalent to a finite-temperature decoding method. Since a lot of computational time is needed for the MPM estimate to calculate the thermal averages, the mean field method, which is a deterministic algorithm, is often utilized to avoid this difficulty. We present a statistical-mechanical analysis of naive mean field approximation in the framework of image restoration. We compare our theoretical results with those of computer simulation, and investigate the potential of naive mean field approximation.

I. INTRODUCTION

In this paper, we investigate the image restoration problem. This problem involves synthesis of an image from a corrupted image with a model of the information available (or assumed) on the source image and the corruption process [1] [2] [3]. The problem lends itself naturally to a Bayesian formulation, which estimates a probability (posterior) for the original image on the basis of the model probability (priors) of the assumed model of the source image and corruption process.

One strategy in image restoration is to use the Bayesian inference by adopting the image that maximizes the posterior probability. This method is called the maximum a posteriori (posterior) probability (MAP) inference. Given a corrupted input image, the MAP inference accepts the image that maximizes the posterior probability as the restored result. The logarithm of posterior probability can be regarded as the energy, so we can consider the MAP inference as an energy minimization problem. Therefore, the image restoration problem can be regarded as an optimization problem. Geman & Geman demonstrated that the MAP inference result can be applied to the image restoration problem by using simulated annealing, which is a tool for searching the ground state [4].

Another strategy is the inference in which the expectation value with respect to the maximized marginal posterior probability at each site in thermal equilibrium is regarded as the original image. This method is called maximizer of the posterior marginals (MPM) inference [5] [6] [7]. In the MPM inference, the posterior probability is given for each set of pixels. In contrast, in the MPM inference, the posterior marginal probability controlled by the temperature $T$ is given for each pixel value. To find a restoration image by the MPM inference, we should calculate thermal average for each pixel value. The MPM inference includes the MAP inference [8] [9] because, at the limit of the temperature $T \to 0$, the MPM inference becomes equivalent to the MAP inference. In general, however, the MPM inference is evaluated at $T > 0$, so this method is also called “finite temperature restoration”, and the MPM inference has an advantage over the MAP inference in the restoration ability of each pixel [9].

Recently, the MPM inference has been discussed in the field of error correcting code [5] [6] [7]. The problem of error correcting code is similar to the image restoration problem in the sense that the received bit sequence, which corresponds to the image represented by a set of pixels, is corrupted by noise, and the receiver tries to retrieve the original bit sequence/image from the noisy one. The major difference between error correcting code and image restoration is that the image restoration problem usually gives only the corrupted image and not other additional redundant information. Therefore, in image restoration, we usually assume alternative information for the original image such as the prior probability. In the field of error correcting code, Rujan has pointed out that carrying out the decoding procedure not at the ground state but at a finite temperature is effective [5]. Sourlas used the Bayes formula to re-derive the finite-temperature decoding of Rujan’s result under more general conditions [10].

Assuming the number of pixels as $N$ and each pixel as a binary unit that can take $\{-1, +1\}$, the number of feasible combinations of pixel values becomes $2^N$. Because of the existence of local minima, finding the ground state task is very difficult with a method such as the gradient decent. Thus, to apply the MAP or MPM inference, we

*Electronic address: shouno@ics.nara-wu.ac.jp*
usually use a relaxation method as typified by annealing technique. Moreover, the MPM inference involves a more important problem as described below. In the MAP inference, once the system reaches the macroscopic equilibrium state, each pixel value can be properly determined with its probability as 1. On the other hand, in the MPM inference, when the system reaches the macroscopic equilibrium state, each pixel value cannot be determined uniquely because the probabilities of each binary state have some finite values in the finite temperature decoding. Therefore, we should calculate thermal averages for each pixel, and this requires many samplings.

In this study, we discuss an approximation that replaces the stochastic dynamics of the MPM inference with deterministic dynamics. In statistical mechanics, this approximation is called the “naive mean field equation (NMFE)” \( \text{[9]} \). The NMFE has usually been applied for emulating the behavior of the stochastic unit. By applying the NMFE, each restoring unit, which is defined as a stochastic binary unit, is replaced by a deterministic analog unit that takes \([-1, +1]\). One important advantage of applying NMFE is the ability to reduce calculation cost by eliminating the need to calculate thermal average, which requires many samplings. The NMFE has been applied to several combinatorial optimization problems such as the traveling salesman problem (TSP) \( \text{[10]} \). However, almost all researchers who have used NMFE have not pursued quantitative issues such as the accuracy of the approximation.

We formulate the image restoration problem using the Bayes inference in the manner of Nishimori & Wong’s formulation \( \text{[11]} \). Recently, from the statistical-mechanical point of view, Nishimori & Wong have formulated the image restoration problem by introducing a mean field model for binary image restoration. They analyzed the model theoretically by the replica method. We applied NMFE to the formulation and analyzed it by the replica method in the manner of Bray et al. \( \text{[9]} \). In \( \text{[11]} \), we compare the results between our analysis and computer simulations. Within the limits of the mean field approximation, our analysis showed agreement with the computer simulations. However, the prior probability derived from the mean field approximation, which is called an infinite range model, is usually not practical for a real image prior. Therefore, we discuss the difference between infinite range model and its origin, called the nearest 0 interaction model.

\[ P_s(\{\xi_i\}) = \frac{1}{Z(\beta_s)} \exp(\beta_s \sum_{i<j} \xi_i \xi_j), \quad (1) \]

\[ Z(\beta_s) = \text{Tr}_\xi \exp(\beta_s \sum_{i<j} \xi_i \xi_j), \quad (2) \]

where \( \beta_s > 0 \). The operator \( \text{Tr}_\xi \) means trace, a sum over all possible \( 2^N \) states of \( \xi \). \( Z(\beta_s) \) indicates the partition function of this prior probability. Assuming \( \beta_s > 0 \), this probability suggests that pixel values \( \xi_i \) and \( \xi_j \) have a tendency to take the same value.

In image restoration, the interaction between each pixel should be described by a local rule, such as a sum of the nearest neighbors. For analysis by the mean field theory, however, we replace this local interaction rule with a global one, that means each pixel has interactions to all other pixels.

In the conventional image restoration framework, the transmitting signal is only the raw code \( \xi \) (upper path in Fig.\( \text{[1]} \)). However, considering the similarity of the formulation of the error-correcting code problem using a Bayesian inference \( \text{[12]} \), it is natural to introduce transmission of a redundant code for better image restoration. Nishimori & Wong proposed the in-
Introduction of a redundant code called the “parity check code” (lower path in Fig. 1) as well as the raw code. For comparison with their results, in this study we also adopted transmission of the parity check code. For the parity check code, we adopted a 2-body interaction term denoted by $\xi_i\xi_j/N$. The reason for dividing $\xi_i\xi_j$ by $N$ is to apply the mean field theory. At the receiver side, the transmitting signal $\xi_i$ corresponds to $\tau_i$, and the parity check code $\xi_i\xi_j/N$ corresponds to $J_{ij}$. The posterior probability based on observation signal $P(\{\xi_i\}|\{J_{ij}\},\{\tau_i\})$ can be represented by the Bayes formula:

$$P(\{\xi_i\}|\{J_{ij}\},\{\tau_i\}) \propto P_{\text{out}}(\{J_{ij}\},\{\tau_i\}|\{\xi_i\})P_s(\{\xi_i\}).$$  

When the receiver can guess the form of the original image posterior probability, the restoring posterior probability can be assumed by replacing the original pixel value $\xi_i$ with the estimated pixel value $\sigma_i$. In general, however, the receiver could not guess the original image prior probability accurately. In order to evaluate the performance of the restoration ability by the difference between the original image prior and the restoring prior, we assume that the restoring image prior probability $P_m(\cdot)$ has a different parameter $\beta_m$ from the original image prior $P_o(\cdot)$.

$$P_m(\{\sigma_i\}) = \frac{1}{Z(\beta_m)} \exp(-\beta_m \sum_{i<j}^N \sigma_i\sigma_j),$$  

$$Z(\beta_m) = \text{Tr}_\sigma \exp(-\beta_m \sum_{i<j}^N \sigma_i\sigma_j),$$

where $\beta_m > 0$. Each form of prior is identical mathematically except for the parameters of interaction strength. We primarily discuss the $\beta_o \neq \beta_m$ case.

$P_{\text{out}}(\{J_{ij}\},\{\tau_i\}|\{\xi_i\})$ is a conditional probability of the observation signal, and it is a probability expression of the corrupting process in the transmission channel. In this study, we assume that the noise added in the corrupting process is similar to Gauss distribution as follows:

$$P_{\text{out}}(\{J_{ij}\},\{\tau_i\}|\{\xi_i\}) \propto \exp\left(-\frac{N}{2J^2} \sum_{i<j} (J_{ij} - \frac{\beta_m}{2} \xi_i\xi_j)^2 - \sum (\tau_i - \tau_0\xi_i)^2 \right).$$  

The first term in the exponential indicates that the noise corresponds to the parity check code, and the second term corresponds to the raw code. For the first term in the exponential, the random variables $J_{ij}$ follow the normal distribution, whose average is $J_0\xi_i\xi_j/N$, and the variance is $J^2/N$. In the second term in the exponential, the random variables $\tau_i$ follow the normal distribution whose average is $\tau_0\xi_i$, and the variance is $\tau^2$.

Nishimori & Wong discussed the macroscopic characteristic of this system by using statistical mechanics [7]. They introduced a system that has the following Hamiltonian:

$$H = -\beta \sum_{i<j} J_{ij}\sigma_i\sigma_j - \frac{\beta_m}{N} \sum_{i<j} \sigma_i\sigma_j - h \sum_i \tau_i\sigma_i,$$  

They calculated the free energy by the replica method for averaging probabilities $P_s(\{\xi_i\})$ and $P_{\text{out}}(\{J_{ij}\},\{\tau_i\}|\{\xi_i\})$. Ignoring the 2-body interaction term, that is $\beta = 0$, this Hamiltonian consists of the prior probability term and the observed data fitting term. This form is equivalent to the conventional cost function form.

In the MAP inference, the estimation value of restored pixels $\sigma$ results in the minimum of the Hamiltonian defined in eq. (8). On the other hand, the MPM inference, which corresponds to the finite temperature estimation, estimates pixel values by the $\text{sgn}(\langle \sigma_i \rangle)$, where $\langle \sigma_i \rangle$ means the thermal average. Considering the limit of the temperature at 0, the MPM inference becomes equivalent to the MAP inference.

We introduced the quantity overlap $M_o$ to evaluate the restoration ability measure for the MAP inference.

$$M_o = \frac{1}{N} \sum_i \xi_i \text{sgn}(\langle \sigma_i \rangle).$$

Nishimori & Wong suggested that the minimization of the Hamiltonian (7) and the maximization of the overlap are not equivalent [7]. Moreover they also indicated that the MPM inference (finite temperature estimation) has better restoration ability than the MAP inference in the sense of the overlap:

$$M_o = \text{Tr}_{\xi_i,\tau_i,\sigma_i} P_s(\{\xi_i\})P_{\text{out}}(\{J_{ij}\},\{\tau_i\}|\{\xi_i\}) \times \xi_i \text{sgn}(\langle \sigma_i \rangle).$$

This overlap can be evaluated by averaging probability variables with the statistical-mechanical technique.

B. Naive mean field approximation

In the finite temperature decoding which corresponds to the MPM inference, we should calculate the thermal
average $\langle \sigma_j \rangle$, which means the estimation value of $\sigma_i$. We evaluated the calculation cost of MPM inference by computer simulation. As we will show that calculating this average requires about 50 times longer computational time than that for achieving to the macroscopic equilibrium state. This result will be given in §III.

In this study, to find the ground state, we introduce the analog neural network model for approximation proposed by Hopfield & Tank [11]. By this approximation, each Ising unit is replaced by an analog unit that can take continuous value $[-1,+1]$, and the output of each analog unit can be regarded as the thermal average of the corresponding Ising spin unit, which can take two states $\sigma_i = \pm 1$. For replacing Ising unit to analog one, we introduced a Hamiltonian as substitution for eq. (6):

$$
\mathcal{H} = M \left[ -\beta \sum_{i<j} J_{ij} \hat{s}_i \hat{s}_j - \frac{\beta_m}{N} \sum_{i<j} \hat{s}_i \hat{s}_j + h \sum_i \tau_i \hat{s}_i - u_i \right],
$$

(10)

where the unit $\hat{s}_i$ is an analog unit. $M$ is a scaling factor described as below. Following the manner of Bray et al. [8], we assumed each $i$th site consists of $M$ Ising units, and the analog unit output $\hat{s}_i$ is calculated by the average of $M$ Ising units $\tilde{\sigma}_{ia}$ (see Fig. 2).

$$
\hat{s}_i = \frac{1}{M} \sum_{a} M \tilde{\sigma}_{ia}.
$$

(11)

Considering the limit $M \to \infty$, each output $\hat{s}_i$ can take an analog value $[-1,+1]$. When $M$ is a finite value, the analog unit output $\hat{s}_i$ is called ‘binominal spins’ which can take $-1, -1 + \frac{2}{M}, \ldots, 1 - \frac{2}{M}, 1$ with binominal distribution. So that, we can define ‘spin weight function’ as

$$
W(\hat{s}_i) = M \text{Tr} \left( M \hat{s}_i - \sum_{a=1}^{M} \tilde{\sigma}_{ia} \right) = M \frac{1}{2\pi} \int_{-\infty}^{+\infty} du_i \exp(M(-u_i \hat{s}_i + \ln 2 \cosh(u_i))).
$$

(12)

The stationary equations are

$$
0 = \beta \sum_{i<j} J_{ij} \hat{s}_i \hat{s}_j + \frac{\beta_m}{N} \sum_{i<j} \hat{s}_i \hat{s}_j + h \sum_i \tau_i \hat{s}_i - u_i,
$$

(14)

$$
0 = -\dot{s}_i + \tanh u_i.
$$

(15)

Eliminating $u_i$, we can obtain

$$
\dot{s}_i = \tanh(\beta \sum_j J_{ij} s_j + \frac{\beta_m}{N} \sum_j \hat{s}_j + h \tau_i).
$$

(16)

From the manner of Hopfield & Tank [11], we can also consider a discrete synchronous updating rule:

$$
\hat{s}_{i}^{t+1} = \tanh(\beta \sum_j J_{ij} s_j^{t} + \frac{\beta_m}{N} \sum_j s_j^{t} + h \tau_i),
$$

(17)

where $s_i^t$ denotes the analog unit output at time $t$. When the system described by eq. (17) reached to the equilibrium state $s_i^\infty$, whole units should satisfy eq. (16). Therefore, to investigate the equilibrium state of dynamics (17), we should take the analog Hamiltonian described by eq. (11) correctly. In the equilibrium state, each analog unit state expressed by $s_i^\infty$ corresponds to $\langle \sigma_i \rangle$, i.e. $s_i^\infty$ can be regarded as the thermal average of $\sigma_i$.

This analog unit replacement is sometimes called the “naive mean field equation (NMFE) approximation”. From eq. (17), this system follows deterministic dynamics, so that it is no need to calculate the thermal average of stochastic unit; therefore, the calculation cost is lower than the stochastic Ising model.

However this replacement from stochastic Ising unit $\sigma_i$ to the deterministic analog unit $s_i$ is an merely substitution. Thus, we should analytically investigate the result of applying NMFE by analog model to compare it with the estimated thermal average of the Ising model. Hence, to investigate the equilibrium state of the system described by eq. (16), we should analyze the system that has the Hamiltonian described by eq. (11).

**1. Replica method**

We now analyze the system that has the Hamiltonian described by eq. (11) by the “replica method”, which is a standard statistical analysis tool. The $n$ replicated partition function can be expressed as:

$$
[Z^n] = \text{Tr} \left[ \prod_{i<j} dJ_{ij} \left( \prod_{i=1}^{N} \frac{d\tau_i}{\sqrt{2\pi\tau^2}} \right) \int d\xi \prod_{i} d\tau_i \right] \times P_s(\{\xi_i\}) \prod_{j} \int d\tau_j \left\{ \sum_{i,a} \sqrt{\text{det}(\tilde{\sigma}_{ia})} \tilde{\sigma}_{ia} \tilde{\sigma}_{ia} \right\} \times \exp \left( \frac{\beta}{M} \sum_{i,j,a,b} J_{ij} \tilde{\sigma}_{ia} \tilde{\sigma}_{jb} \right) \times \exp \left( \frac{\beta_m}{MN} \sum_{i,j,a,b} \tilde{\sigma}_{ia} \tilde{\sigma}_{ib} + \sum_{i,a} h \tau_i \tilde{\sigma}_{ia} \right),
$$

(18)
where operator $\text{Tr}$ means the sum over all states about $\{\sigma_i^m\}$ and $\{\xi\}$. We analyzed this replicated partition function by the standard replica method. The replica symmetry solution can be described as:

$$m_0 = \tanh(\beta_s m_0),$$  
$$m = \frac{\text{Tr}_{\xi} e^{\beta_s m_0 \xi} \int Dx \hat{F}(U(x))}{2 \cosh(\beta_s m_0)},$$  
$$t = \frac{\text{Tr}_{\xi} e^{\beta_s m_0 \xi} \int Dx \hat{F}(U(x))}{2 \cosh(\beta_s m_0)},$$  
$$q = \frac{\text{Tr}_{\xi} e^{\beta_s m_0 \xi} \int Dx \hat{F}(U(x))^2}{2 \cosh(\beta_s m_0)},$$  
$$\chi = \frac{1}{\sqrt{\hbar^2 \tau^2 + \beta^2 J^2 q}} \frac{\text{Tr}_{\xi} e^{\beta_s m_0 \xi} \int Dx x \hat{F}(U(x))}{2 \cosh(\beta_s m_0)},$$

where $U(\cdot)$ means

$$U(x) = \sqrt{\hbar^2 \tau^2 + \beta^2 J^2 q x + \beta_s m + (\hbar \tau_0 + \beta J_0) \xi},$$

and the function $\hat{F}(\cdot)$ is a solution of the self consistent equation:

$$F(x) = \tanh(U(x) + \beta^2 J^2 \chi F(x)).$$

Using these solutions, we could obtain the overlap $M_o$:

$$M_o = \frac{1}{N} \sum_i \xi_i \text{sgn}(\xi_i) = \frac{\text{Tr}_{\xi} e^{\beta_s m_0 \xi} \int Dx \xi_s \text{sgn}(U(x))}{2 \cosh(\beta_s m_0)}.$$  

### III. RESULTS

In this section, we compare the analysis results between applying the NMFE approximation (analog model) and estimated thermal average of the Ising model. We also compare these theoretical analyses with the computer simulation results.

In the following subsection from III A to III C, we treated infinite range Ising/analog model. In the infinite range model, the original image prior probability is defined by eq. (1), and the image is corrupted by the noisy channel according to eq. (3). In the infinite range Ising model, we assumed restoring prior probability as eq. (1), and the Hamiltonian could be derived as eq. (3). On the other hand, in the infinite range analog model, we defined the Hamiltonian as eq. (2).

However, for image prior, the infinite range interaction is just a little strange assumption. Thus, we introduced models whose prior have only nearest neighbor interaction in III D. We call them as Ising model with nearest neighbor interactions and analog model with nearest neighbor interactions respectively. Unfortunately, it is difficult to treat the nearest neighbor interaction models analytically. Therefore, we calculated the result of nearest neighbor interaction models by computer simulation and compared them with the result of the infinite range models. In the nearest neighbor interaction models, we assume that the corruption process in transmitting is same as infinite range model, that is defined by eq. (4).

In the Ising model Monte Carlo simulation, we use the asynchronous Glauber dynamics, i.e. we selected one site
σ_i randomly, and calculate probability as

\[ P(\sigma_i = \pm 1) = \frac{(1 \pm \tanh(\beta \sum_j J_{ij} \sigma_j + \frac{\beta}{N} \sum_j \sigma_j + h \tau_i))}{2} \]

According to the probability the selected site \( \sigma_i \) is decided.

On the other hand, for analog simulation, we used discrete synchronous update rule described as eq. (17), i.e. whole units are updated simultaneously. We consider each unit value in the equilibrium state as its thermal average.

### A. Without parity check code

First, we discuss the case where no parity check code is used (\( \beta = 0 \)). This case corresponds to the conventional image restoration problem. In the analysis, when \( \beta = 0 \), the order parameter equations obtained in the previous section are identical to the result of the Ising analysis suggested by Nishimori & Wong [7]. Thus, the MPM inference by the stochastic infinite range Ising model and deterministic infinite range analog model are equivalent. Figure 3 illustrates the restoring ability of finite temperature decoding using the infinite range Ising stochastic units. Figure 4 shows the restoring ability using infinite range analog model. In each figure, the vertical axis means the overlap \( M_o \), and the horizontal axis means the restoring temperature \( T_m \), which is corresponds to the restoring image prior coefficients \( \beta_m \) (\( T_m = \beta_m^{-1} \)). The solid line means the analysis result by the replica method, and each error bar shows the quartile deviation obtained by the computer simulation.

We set the parameter \( T_s \), which is the reciprocal of the original image prior coefficients \( \beta_s \) (\( T_s = \beta_s^{-1} \)), at \( T_s = 0.9 \). At the limit of \( T_m \to 0 \), the MPM inference becomes equivalent to the MAP inference. In both figures

### B. Convergence speed

The biggest advantage of the analog model over the Ising model is its lower calculation cost. In this section, we discuss the calculation time that should be set in the computer simulation for each method.

We observed the convergence time of the overlap in each simulation. Figure 5 shows the convergence speed.
In the infinite range Ising model, we adopt the synchronous update and define $N_m$ updates as one Monte-Carlo step (1 MCS). $N_m$ means the number of units in the simulation. A synchronous analog update corresponds to 1 MCS in the sense of the number of update units. The horizontal axis of Fig. 8 indicates the calculation time measured by MCS. The vertical axis means the average of overlap at that time. Each model achieved the macroscopic equilibrium state in 20 MCS. The analog model adopts a deterministic algorithm, we should calculate the thermal average of $s_i$ by sampling. Figure 8 shows that needs about 1000 MCS to converge. In this result, the deterministic infinite range analog model converged 50 times faster than the stochastic infinite range Ising model in finite temperature decoding.

C. Using parity check code

In the replica method analysis, the infinite range analog model and the infinite range Ising model are equivalent in the case using $\beta = 0$. This is derived from the fact that the susceptibility $\chi$ remained at 0 under $\beta = 0$. Using the parity check code for better restoration, the analysis results differ. In the case of $\beta > 0$, $\chi$ becomes greater than 0, and the term corresponding to $\chi$ in eq. 23 is effective. This term is called the Onsager reaction term in statistical mechanics. As a result, a difference in analysis between the Ising model and the infinite range analog model appears.

When the noise variance of the parity check code is small (e.g., $J \sim 0.6$), the infinite range analog model and the infinite range Ising model are almost the same in the sense of overlap. In this case, little noise exists in the parity check code, and better image restoration is possible. Figure 9 shows a comparison of overlaps obtained by each replica analysis. The horizontal axis indicates the parameter $\beta$, and the vertical axis indicates the overlap $M_o$. We assumed that the decoding temperature is optimal ($T_m = T_s (= 0.9)$) and that the parameter is $h = 1$. Comparing $\beta$ with $h$, $\beta$ becomes large, the restoring system tends to rely on the parity check rather than the raw code. Each restoration result appears the same, and better restoration is achieved than in the case transmitting the raw code only.

Figure 8 shows the case of large noise variance in the parity check code channel ($J = 1.0$). The horizontal axis signifies $\beta$. In this case, each method restores better than when only the raw code channel is used. However, relying on $\beta$ too much results in poorer restoration result than when only the raw code channel is used. A difference between the infinite range analog model and the infinite range Ising model appears as restoration worsens. When the noise variance of the parity check code becomes large, the restoration ability of infinite range analog model is not as good as that of the infinite range Ising model.

Figure 9 shows the computer simulation result of the infinite range Ising model under the noise variance parameter $J = 1.0$, and Fig. 8 shows the computer simulation result of the infinite range analog model. Each simulation matches the analysis curve, and the performance of infinite range analog model simulation is not much better than that of the infinite range Ising simulation. However, when we properly set the ratio of the coefficients $\beta, \beta_m, \text{and } h$, the performance of each is optimized and the peak of ability for each is almost the same. Thus when we can estimate the hyper-parameters properly, there is no difference between these two methods.

D. Ising/analog model with nearest neighbor interaction

In each infinite range model we compared, we assumed that all units have interactions with each other in prior
probabilities to use the mean field approximation. However, it is more natural that these probabilities are defined by their nearest-neighbor interactions. We call the model whose encoding/decoding prior probability is defined by the nearest neighbor interaction on the pixel lattice the Ising/analog model with nearest neighbor interaction. It is difficult to analyze the nearest neighbor interaction model with a statistical tool because of its local interactions. Nishimori & Wong compared the analysis result of the Ising model with nearest neighbor interaction solved by computer simulation [7]. Figure 10 shows the results of our reproduction. The image size is 400 × 400 pixels, and the original image’s prior probability is denoted:

\[ P_s (\{\xi_i\}) = \frac{1}{Z(\beta_s)} \exp(\beta_s \sum \langle ij \rangle \xi_i \xi_j), \quad (27) \]

\[ Z(\beta_s) = \text{Tr} \sum (\xi_i \exp(\beta_s \sum \langle ij \rangle \xi_i \xi_j)), \quad (28) \]

where the summation \(\sum \langle ij \rangle\) denotes a sum extending over all pairs of neighboring sites on the pixel lattice.

In this simulation, we adopt \( T_s = 1/\beta_s = 2.15 \), generate the image \(\{\xi_i\}\) by the prior probability (27), and corrupt by flipping each pixel by the probability \(p = 0.1\). We use only the raw code channel, not the parity check code channel in this simulation. By using the mean field theory, Nishimori & Wong proved that the overlap \(M_o\) is maximized despite of the dimensions of the system (image size), when the decoding temperature \(T_m\) equals the image generating temperature \(T_s\). In the simulation, we set \( T_s = T_m = 2.15 \). Figure 11 shows the reproducing result of the Ising/analog model with nearest neighbor interaction obtained by the computer simulation. With the nearest-neighbor interaction model, it is clear that the overlap of the analog model is smaller than that of the Ising model. We believe that this phenomenon occurred because the analog model does not gather in the effect of fluctuation by other pixel values.

Thus, we should consider the nearest neighbor interaction prior effect in the computer simulation. Figure 11 shows an example of the computer simulation result. Each image has an area of 400 × 400 pixels. Figure 11(a) shows the original image. In this image, noise-like patterns are inherently contained. Figure 11(b) shows the corrupted image. Figure 11(c) shows the restoration result by the Ising model, and Figure 11(d) shows the result by the analog model. The overlap of the Ising model (Figure 11(c)) becomes \(M_o = 0.879\), and the analog model (Figure 11(c)) becomes \(M_o = 0.837\). In addition, when we applied the image restoration by MAP inference, the noise-like patterns contained in the original image (Figure 11(a)) were eliminated by the prior probability term. Thus, to maintain the fine structure of the original image, the MPM inference is an effective method. This characteristics was confirmed in the analog model by the computer simulation.

IV. DISCUSSION

In this research, we discussed the image restoration problem by finite temperature decoding. In the MPM inference proposed by Nishimori & Wong, many trials are required to calculate the thermal average for each unit. We tried to replace this operation with deterministic dynamics called naive mean field approximation. At first, we applied mean field approximation to the infinite range prior model and analyzed this approximated system by the replica method. We also quantitatively compared the restoration ability of the Ising stochastic model with that of the analog deterministic model. As a result, we proved that their abilities are equivalent when we used only the raw code signal, that is, equivalent to conventional image restoration. Assuming a redundant signal, called the parity check code, the result of each method differed slightly. However, when we adopted the ratio of the coefficients properly, that is, the coefficients \(\beta, h,\) and \(\beta_m\), the restoring ability of each method showed no difference. This means that if we can estimate these coefficients, called hyper-parameters, we can apply the analog deterministic method without worrying about the restoration ability. Moreover, the analog deterministic method is 50 times faster than the Ising stochastic method.

In the future problem, In analog model, Horiguchi pointed out that the inclination of output function around ±1 cause constitutive difference [11]. Thus, we should consider this effect.

Moreover, we will attempt to propose a more effective analog approximation method for the nearest neighbor interaction model.
FIG. 11: Example of image restoration. The original image was created with a Ising/analog model with nearest neighbor prior model. (a) Original image; (b) Noised image; (c) Ising restoration result by MPM estimation; and (d) Analog restoration result by MPM estimation.