KCNE1 rs1805127 Polymorphism Increases the Risk of Atrial Fibrillation: A Meta-Analysis of 10 Studies

Chang Liang1, 3, Xiankai Li1, * , Yawei Xu1, Qingyong Chen1, Yadong Wu1, Wan Wang2, Weiming Li1, * , Mantang Qiu3, 4

1 Department of Cardiology, Shanghai Tenth People’s Hospital, Tongji University, Shanghai, China, 2 The Key Laboratory of Cardiovascular Disease and Molecular Intervention of Nanjing Medical University, Nanjing, China, 3 The Fourth Clinical College of Nanjing Medical University, Nanjing, China, 4 Department of Thoracic Surgery, Cancer Institute of Jiangsu Province, Nanjing, China

Abstract

Background: Atrial fibrillation (AF) is one of the most common types of arrhythmia in humans. Recently, many studies have investigated the relationship between human atrial fibrillation and the single nucleotide polymorphism (SNP) of rs1805127 (A>G) in KCNE1 gene, but the results were still inconsistent and inconclusive.

Method: Electronic databases and bibliographies of retrieved studies were searched. We performed a meta-analysis of ten case-control studies, including 2099 cases and 2252 controls, to evaluate the association of rs1805127 polymorphism (A>G) with the risk of AF. Random-effects model was used when the heterogeneity was obvious; otherwise, fixed-effects model was applied. Meta-regression was performed to examine potential source of heterogeneity. Egger’s test and Begg’s test were used to detect publication biases.

Results: The results showed a significantly increased risk of AF in homoyzgote comparison (GG vs. AA: OR = 1.899, 95%CI: 1.568, 2.300; P heterogeneity = 0.217), heterozygote comparison (GA vs. AA: OR = 1.436, 95% CI:1.190, 1.732; P heterogeneity = 0.739), dominant model(GA /GG vs. AA: OR =1.624, 95%CI: 1.361, 1.938; P heterogeneity =0.778) and recessive model (GG vs. GA/AA: OR = 1.394, 95%CI:1.152, 1.686; P heterogeneity =0.03). Meta-regression revealed that the sample size and the types of AF were the source of the heterogeneity.

Conclusion: The rs1805127 polymorphism (A>G) of KCNE1 is associated with an increased risk of AF, which suggests the rs1805127 polymorphism of KCNE1 gene may play an important role in the pathogenesis of AF.

Introduction

Atrial fibrillation (AF) is the most common type of sustained tachyarrhythmia in human, and the prevalence increased year by year. However, the pathogenesis of AF remains unclear. Currently, a number of studies have shown that mutations in ion channel genes may be the risk factors of AF [1,2].

The KCNE1 gene was first discovered by Murai in 1989[3], which located in 21q22.1-22.2. KCNE1 encodes the β-subunits of the delayed rectifier potassium current channel (IKs) in human heart, which is also called Mink protein, including 130 amino acids. The functional rs1805127 polymorphism (A>G) of KCNE1 gene leads to a serine to glycine substitution [4]. Over the last decade, a number of case-control studies have been conducted to investigate the association between rs1805127 polymorphism and AF. Lai reported that the variant G allele was more common in AF patients compared with that of control subjects, the result showed that KCNE1 polymorphism was associated with atrial fibrillation [5]. This conclusion was supported by the studies reported by Prystupa [6] and Fatini [7]. However, this association was not observed in other studies [8–11]. During the past two years, three case-control studies [12–14] with larger sample size were performed, but the result remain inconsistent and inconclusive.

To determine the association between KCNE1 rs1805127 polymorphism and the risk of AF, we performed this meta-analysis of published case-control studies to attain better results and hence more details and accurate risk estimation.

Methods

We done this meta-analysis according to the PRISMA guideline (Supporting information: Table S1). Appropriate case-control studies were extracted by electronic search of databases and manual search of references of relative articles and reviews. We searched PubMed/MEDLINE, EMBASE, and China National Knowledge Infrastructure (CNKI) using key words of “KCNE1 or potassium voltage-gated channel, Iks-related family, member 1”, “polymorphism or variant” and “atrial fibrillation or AF”. 
There were no limitation of research data and the last research was performed on 8 January, 2012. References of related studies and reviews were manually searched for additional studies.

**Inclusion and Exclusion Criteria**

Inclusion criteria were as follows: (1) case-control studies; (2) investigating the association between the KCNE1 rs1805127 (A>G) polymorphism and risk of AF; (3) studies that provide detailed genotypes data (AA, GA and GG) in AF cases and control groups. Exclusion criteria were as follows: (1) duplicated studies; (2) no controls; (3) no detail genotype frequencies. We did not consider unpublished reports, abstracts, comments, reviews, or editorials.

Two reviewers extracted eligible studies independently according to the inclusion criteria. Disagreements between two reviewers were discussed with another reviewer till consensus was achieved.

**Data Extraction**

Data extraction was conducted independently by two authors with disagreements resolved by discussion. The following data were extracted from eligible study: the first authors’ name, publication year, country where the study was conducted, genotyping methods, ethnicity, source of control, Hardy-Weinberg equilibrium (HWE), number of cases and controls, genotype frequency in cases and controls. Different ethnicities were divided

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**Table 1. Eligible Studies and data.**

| first author      | year | country | ethnicity | source of control | Cases | Controls |
|-------------------|------|---------|-----------|-------------------|-------|----------|
|                   |      |         |           |                   | GG    | GA       | AA   | GG    | GA   | AA   |
| Lingping Lai [9]  | 2002 | China   | Asian     | HB                | 64    | 37       | 7    | 46    | 44   | 18   |
| Ni Aizhen [10]    | 2004 | China   | Asian     | PB                | 54    | 37       | 3    | 72    | 48   | 10   |
| Cinzia Fatini [5] | 2006 | Italy   | Caucasian | HB                | 118   | 155      | 58   | 116   | 207  | 118  |
| Andrzej pryta [6] | 2006 | Poland  | Caucasian | PB                | 24    | 38       | 7    | 3     | 45   | 13   |
| LOU Sheng [7]     | 2006 | China   | Asian     | PB                | 63    | 41       | 7    | 60    | 29   | 12   |
| ZENG Zhi-yu [11]  | 2007 | China   | Asian     | PB                | 71    | 60       | 10   | 55    | 54   | 11   |
| Xu Lixin [8]      | 2008 | China   | Asian     | HB                | 77    | 61       | 9    | 75    | 56   | 16   |
| Yao Juan [15]     | 2011 | China   | Asian     | HB                | 158   | 117      | 28   | 129   | 159  | 40   |
| Yao Juan [16]     | 2012 | China   | Asian     | HB                | 133   | 138      | 36   | 118   | 148  | 64   |
| Mao Ting [17]     | 2012 | China   | Asian     | HB                | 218   | 201      | 69   | 188   | 215  | 85   |

PB: population-based; HB: hospital-based.

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Table 2. Meta-analysis results.

|                | GG vs. AA | GA vs. AA | GA/GG vs. AA | GG vs. GA/AA |
|----------------|-----------|-----------|--------------|--------------|
|                | N         | OR        | P_h          | OR           | P_h          | OR        | P_h          | OR           | P_h          |
| Total          | 10        | 1.899(1.568,2.300)* | 0.217         | 1.436(1.190,1.732)* | 0.739         | 1.624(1.361,1.938)* | 0.778         | 1.394(1.152,1.686) * | 0.03         |
| Ethnicities    |           |           |              |              |              |           |              |              |              |
| Asian          | 8         | 1.756(1.407,2.191) * | 0.768         | 1.401(1.122,1.748)* | 0.561         | 1.564(1.270,1.926)* | 0.690         | 1.328(1.160,1.521) * | 0.344        |
| Caucasian      | 2         | 4.822(0.711,32.679) | 0.107         | 1.529(1.074,2.177)* | 0.958         | 1.785(1.280,2.490)* | 0.536         | 3.615(0.563,23.228) | 0.004        |
| Source of control |         |           |              |              |              |           |              |              |              |
| HB             | 7         | 1.831(1.491,2.248)* | 0.532         | 1.393(1.140,1.703)* | 0.535         | 1.586(1.314,1.916)* | 0.690         | 1.433(1.250,1.641) * | 0.419        |
| PB             | 3         | 2.724(1.111,6.679)* | 0.067         | 1.757(1.040,2.967)* | 0.731         | 1.917(1.160,3.167)* | 0.792         | 1.482(0.779,2.820) | 0.006        |
| Sample size    |           |           |              |              |              |           |              |              |              |
| Large study    | 4         | 1.764(1.420,2.192)* | 0.145         | 1.338(1.083,1.653) * | 0.919         | 1.525(1.251,1.859)* | 0.889         | 1.445(1.246,1.675)* | 0.581        |
| Small study    | 6         | 2.460(1.628,3.719)* | 0.552         | 1.862(1.236,2.805)* | 0.447         | 2.062(1.396,3.047)* | 0.534         | 1.404(0.927,2.127) | 0.008        |

num: number of studies included; OR: odds ratio; P_h: p value for heterogeneity; PB: population-based; HB: hospital-based; *OR with statistical significance; large study: studies with more than 300 participants; small study: studies with less than 300 participants.
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Figure 2. Forest plot of homozygous comparison (GG vs. AA) for overall comparison in fixed-effects model. (OR = 1.899, 95%CI: 1.568, 2.300; P_heterogeneity = 0.217).
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into Asian and Caucasian. Source of control in the studies were
defined as hospital-based (HB) and population-based (PB). An
online program (http://ihg.gsf.de/cgi-bin/hw/hwa1.p1) was used
to test the Hardy-Weinberg equilibrium in the controls [15], and a
p<0.05 indicates disequilibrium of HWE.

Statistical Analysis

Odds ratios (OR) and 95% confidence intervals (CI) were used
to measure the association strength between the KCNE1
rs1805127 polymorphism and AF risk. In this meta-analysis, we
evaluated the overall AF risk in four comparison models:
homozygote comparison (GG vs. AA), homozygote comparison
(GA vs. AA), recessive model (GG vs. GA/AA) and dominant
model (GG/GA vs. AA). Subgroup analyses were carried out to
explore the confounding factors: ethnicities, source of control and
sample size. Sensitivity analyses were performed to identify
individual study effect on pooled results and test the reliability of
results.

The Cochran’s Q statistic was used to evaluate the heteroge-
nity between studies [16]. When the P value was <0.10, the
heterogeneity were considered as significant, and the pooled AF
risk was estimated by the random-effects model; otherwise, fixed-
effects model was applied [17]. Meta-regression analysis was
performed to find the source of heterogeneity, and a p<0.05 was
considered significant [18].

We used visual inspection of asymmetry in funnel plots, Begg’s
test and Egger’s test to statistically potential publication bias, and a
P<0.05 was considered significant [19]. All statistical analyses
were calculated with STATA software (StataCorp, College
Station, TX). And all P values were two-side with a significance
level of 0.05.

Results

Characteristics of eligible studies

Finally, 10 studies investigating the association between the
rs1805127 polymorphism of KCNE1 gene and AF were identified
according to the inclusion and exclusion criteria, including 2099
cases and 2252 controls [4–6,8–14]. The detailed screening
process was showed in Figure 1, and the characteristics of eligible
studies are listed in Table 1. Blood sample was used for genotyping
in all studies. Among the ten studies, eight were performed in
Asian; two were performed in Caucasian. The sample size of four
of the ten is larger than 300 (classified as “large”), and six of the
ten were less than 300 (classified as small). Most of the studied are
Hospital-Based (HB, seven studies), and the others were Popula-
tion-Based (PB). HWE of genotype distribution in the controls was tested by the online program and the genotypes distribution in controls was not in agreement with HWE in two studies [8,9]. Since no genotype error was detected, these two studies were not excluded.

Result of meta-analysis

In overall analysis, we found rs1805127 polymorphism was associated with increased risk of AF in all four comparisons: homozygote comparison (GG vs. AA: OR = 1.899, 95%CI: 1.568, 2.300; \( P_{\text{heterogeneity}} = 0.217 \), Table 2, Figure 2), heterozygote comparison (GA vs. AA: OR = 1.436, 95%CI: 1.190, 1.732; \( P_{\text{heterogeneity}} = 0.739 \), Table 2, Figure 3), dominant model (GA/GG vs. AA: OR = 1.624, 95%CI: 1.361, 1.938; \( P_{\text{heterogeneity}} = 0.778 \), Table 2, Figure 4) and recessive model (GG vs. GA/AA: OR = 1.394, 95%CI: 1.152, 1.686; \( P_{\text{heterogeneity}} = 0.03 \), Table 2, Figure 5).

Subgroup analyses for ethnicity suggested that the rs1805127 polymorphism was associated with an increased risk of AF among Asian population (Table 2). But for Caucasian population, the increased risk of AF was observed in heterozygote comparison and dominant model, but not in homozygous comparison and recessive model (Table 2). Analyses by source of control showed that there was significant association between the rs1805127 polymorphism of KCNE1 and increased risk of AF both in HB and PB studies (Table 2). Stratified analyses by sample size also suggested that the rs1805127 polymorphism (A>G) of KCNE1 increased risk of AF both in large studies and small studies (Table 2).

Heterogeneity and Sensitivity Analysis

The result of heterogeneity between included studies were shown in Table 2. Heterogeneity was significant in recessive model of overall analysis. We performed meta-regression to detect source of heterogeneity and sensitivity analysis to explore each study's influence on the total results. Meta-regression revealed that source of control (\( P = 0.047 \)) and ethnicity (\( P = 0.035 \)) but not sample size (\( P = 0.337 \)) were the sources of heterogeneity, which was inconsistent with sub-group analyses results (Table 2). Sensitivity analysis suggested that the study published by Prytupa [6] was responsible for the heterogeneity. After removing this study, the results did not change significantly (OR = 1.362, 95%CI: 1.203, 1.541), and the no significant heterogeneity was observed (\( P = 0.369 \)).
We assessed publication bias by Begg’s funnel plot and Egger’s test. The results showed no evidence of publication bias (homozygote comparison: $P_{\text{Begg}} = 0.097$; dominant model: $P_{\text{Begg}} = 0.052$, recessive model GG vs. GA/AA: $P_{\text{Begg}} = 0.084$), except for heterozygote comparison (GA vs. AA: $P_{\text{Begg}} = 0.049$, Figure 6), further study suggest that the study of Mao [14] was responsible for the asymmetry of funnel plot; we performed trim and fill method. When this title was deleted, there was no evidence of publication bias (GA vs. AA: $P_{\text{Begg}} = 0.128$, Figure 7), and the pooled OR was still significant ($OR = 1.547$, 95% CI: 1.244, 1.924, $P_{\text{heterogeneity}} = 0.834$).

Discussion

KCNE1 gene encodes the β-subunits of the delayed rectifier potassium current channel (IKs) in human heart. The delayed rectifier potassium current channel (IKs) is important for cardiac repolarization, especially at the late phase of phase 3 of action potential, and it’s closely related to action potential duration [20–22]. Recently, many studies have investigated the association between IKs function and heart arrhythmia, mostly in long Q-T syndrome and AF [21,23–26]. Ehrlich and colleagues demonstrated that the rs1805127 G-allele carriers have a higher prevalence of tachyarrhythmia resulting from decreased IKs current, prolonged action potential duration, prolonged relative refractory period, inducing early afterdepolarization under specific conditions [19]. The study reported by Chevillard confirmed that
the KCNE1 gene was over-expressed in atrial tissue in patients with permanent atrial fibrillation [24]. It was also found that the protein KCNE1 gene encoded may work through interacting with other proteins, forming arrhythmic substrate and resulting in atrial fibrillation and maintaining [21]. These results show that KCNE1 gene is important in the regulation of heart rhythm.

In this meta-analysis, we analyzed 10 studies including 2099 cases and 2252 controls and found a significant association between the rs1805127 polymorphism (A>G) of KCNE1 and increased risk of AF in all comparison models of overall analysis.

In subgroup analyses of ethnicity, the significant association was observed in all comparisons in Asian population, but not all comparisons in Caucasian. Small sample size and limited number of studies of Caucasian in this meta-analysis (only two studies available [6,7]) maybe the major reasons of these discrepancies. Additionally, the heterogeneity between two studies was significant (Table 2), which may be also contributed by limited number of studies. The different AF risks in Asians and Caucasians was also reported in other meta-analysis [27]. Additional studies with larger sample sizes will be necessary to clarify this finding.

Subgroup analysis of source of control also showed a positive association, except the recessive model, in PB studies. However, studies of PB were not all complied with HWE[4,9], and other heart disease may be an incentive of AF, like coronary heart disease and valvular heart disease, thus, studies of PB control could not match all high risk factors of AF between case group and control group. Consequently, setting HB persons as a control group is more representative for the result and encounters less bias, which matched cases and controls according to various risk factors.

Since both results from PB and HB suggested that the rs1805127 polymorphism increased risk of AF, this conclusion is reliable.

For heterogeneity, in overall analysis, the results showed no significant heterogeneity except in recessive model. Indicated by the results of sensitivity analysis of recessive model, the study by Prytupa [6] was considered as the source of heterogeneity. Firstly, sample size of Prytupa’s study was small, with only 69 cases and 60 controls. Secondly, the case group of Prytupa’s study was based on lone paroxysmal AF and patients with permanent AF were excluded, which was different from other studies. However, this kind of heterogeneity was difficult to avoid because enough sample of kinds of AF types was difficult. Additionally, meta-regression revealed that source of control and ethnicity but not sample size contributed to heterogeneity.

In this meta-analysis, we analyzed all eligible studies about the association between rs1805127 polymorphism (A>G) of KCNE1 gene and risk of AF, including 2099 cases and 2252 controls. There was no limitation of languages when searching, and the results showed a low chance of publication bias. Some limitation of our meta-analysis should be considered. Firstly, AF was considered as a multi-factorial disease, to achieve more precise estimate, individual data should be adjusted according to some other variables that related to risk of AF, such as family history and coronary heart disease. Secondly, as reported by Fatini [7] and Xu [10], whether there was a joint interaction between rs1805127 polymorphism of KCNE1 and polymorphisms of other genes (such as eNOS T796G gene polymorphism) in the development of AF is unknown. Further experiment studies are needed to find the mechanism of the combined effects. Thirdly, the AF types were not fully specified in studies, uniform adjustment is needed for the diagnosis of AF in each study.

In conclusion, our meta-analysis demonstrates that the rs1805127 polymorphism (A>G) of KCNE1 gene is associated with increased risk of AF; this finding suggests that the rs1805127 polymorphism of KCNE1 gene may play an important role in the pathogenesis of AF. To confirm this association, further experiment studies are warranted.

Supporting Information

Table S1 PRISMA checklist.

Author Contributions

Conceived and designed the experiments: CL XL YX QC YW WL MQ. Performed the experiments: CL XL WL MQ. Analyzed the data: CL XL WW WL MQ. Contributed reagents/materials/analysis tools: CL XL YX QC YW WW WL MQ. Wrote the paper: CL XL MQ.

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