Testing the nexus between stock market returns and inflation in Nigeria: Does the effect of COVID-19 pandemic matter?

Gylych Jelilov | Paul Terhemba Iorember | Ojonugwa Usman | Paul M. Yua

1Department of Economics, Nile University of Nigeria, Abuja, Nigeria
2Faculty of Social Sciences, Nile University of Nigeria, Abuja, Nigeria
3School of Business Education, Federal College of Education (Technical), Potiskum, Nigeria
4Graduated from Nnamdi Azikiwe University, Awka, Nigeria

Correspondence
Ojonugwa Usman, School of Business Education, Federal College of Education (Technical), Potiskum, Yobe State, Nigeria. Email: usmanojonugwa@gmail.com

1INTRODUCTION

The outbreak of the novel coronavirus (SARS-CoV-2) code named COVID-19, which started in Wuhan, the capital of Hubei Province in the People’s Republic of China around December 2019 has become a global pandemic, expanding rapidly across the planet with infection cases and confirmed death rates nearing five million and four hundred thousand respectively as at May 18, 2020 (World Health Organization (WHO), 2020). Apart from the palpable fear that the COVID-19 pandemic has created around the world, it has disrupted global economic activities in a manner that has never been seen since the 1918 influenza pandemic. Over a third of the global population is currently under some forms of lockdown, border closures, and extensive flight cancellations creating extreme disruptions to business and tourism (Ayittey, Ayittey, Chiwero, Kamasah, & Dzuvor, 2020). The economic consequences of the novel coronavirus include interruptions to global supply and demand chains and corresponding local food shortages, depressed asset prices, and extreme uncertainty for both multinational and local businesses (Yousef, 2020).

Of particular importance regarding the economic impact of COVID-19 is the effect of COVID-19 pandemic on stock market volatility in the presence of bubble prices (inflation). This is because, as the uncertainties triggered by the pandemic continues, the effectiveness of policy measures coupled with negative investors sentiments and rising prices are likely to affect long-term investment decisions. Regarding inflationary pressures, however, the Fisher (1930) hypothesis otherwise known as the “Fisher effect” states that the expected nominal stock market returns should be equal to expected inflation plus the real rate of return. In effect, the hypothesis elucidates that equity stocks represent claims against real assets of a business; and as such, may serve as a hedge against inflation (Sokpo, Iorember, & Usar, 2018). If this assertion holds, then investors could sell their financial assets in exchange for real assets when the expected rise in inflation is pronounced. In such a situation, Ioannides, Katrakilidis, and Lake (2005) posits that nominal stock prices should fully reflect...
expected inflation and the relationship between these two variables should be positively correlated ex-ante. In the view of Lawal (2016), the argument of stock market returns serving as a hedge against inflation may also mean that investors are fully compensated for the rise in the general price level through corresponding increases in nominal stock market returns and thus, the real returns remain unchanged.

However, it was not too long that divergent opinions emerged challenging the Fisher hypothesis. In the 1970s, investors found that the Fisher theory faltered in the short and intermediate terms, as stock market returns were negatively related to inflation (Sharpe, 2002). Also, empirical studies by Bodie (1976), Nelson (1976), Fama and Schwert (1997), and Modigliani and Cohn (1979) all adduced empirical evidence to show an inverse relationship between stock market returns and inflation. In 1981, Fama came up with a counter-theory known as the Proxy Hypothesis. The Proxy hypothesis commonly known as the generalized Fama proxy hypothesis argued that the assertion of a positive relationship between stock market returns and inflation was spurious, suggesting that stock market returns do not hedge against inflation.

The literature on the nexus between stock market returns and inflation is well documented. Most recently, Alqaralleh (2020) in a study on the nexus between stock market returns and inflation for the G7 countries; revisited evidence based on nonlinear ARDL, established that expansionary period increase stock market returns more than contractionary period, thereby supporting the Fisher effect. Also supporting the Fisher hypothesis, Kwofie and Ansa (2018) examined the effect of inflation on stock market returns in Ghana and founded a positive long-run relationship between inflation and stock market returns. Other studies that support the Fisher effect include Li, Balcilar, Gupta, and Chang (2016) for China and India; Kumuduni and Jameel (2016) for Colombia; Omair, Yuksel, and Yuksel (2015) for a sample of 52 countries; Mireku, Sarkodie, and Poku (2013) for Ghana; Acaravci, Acaravci, and Ozturk (2011) for Turkey; and Omotor (2010), Ibrahim and Agbeje (2013) and Olufisayo (2013) for Nigeria.

Regarding the generalized Fama proxy hypothesis on the assertion that stock market returns do not hedge against inflation, Singh and Padmakumari (2020) examine the reaction of stock market returns to the inflation announcement using time series data for India from 2012 to 2018 and found an inverse relationship between stock market returns and inflation. Similarly, other empirical studies that affirmed the Fama proxy hypothesis include Sokpo et al. (2018), Emeka and Aham (2016), Uwubanmwen and Eghosa (2015) for Nigeria; Otieno, Ngugi, and Wawire (2017) for Kenya; Silva (2016) for Sri Lanka; Adusei (2014) for Ghana; and Yeh and Chi (2009) for OCED countries.

Interestingly, none of these studies controlled for any kind of pneumatic diseases such as the novel COVID-19 pandemic or any other pandemic in recent history even though, a pandemic disease creates global health and economic crisis with disruptions on both supply and demand, the collapse of oil price, drop in the country’s real GDP, causing panic trading and triggering market-wide circuit breakers.

Arising from the above and given the palpable fear generated by the threat of COVID-19 pandemic, this study represents one of the first efforts towards testing the effect of COVID-19 on the stock market returns-inflation relationship. That is, to determine whether stock market returns hedge against inflation or not in the era of the COVID-19 pandemic? Specifically, the study aims to revisit the stock market returns— inflation relationship, controlling for the effect of COVID-19 pandemic in Nigeria using the Generalized Autoregressive Conditional Heteroskedasticity type models (Standard GARCH and GJR-GARCH), and the Accounting Innovation tests.

The rest of the paper is structured as follows: Section 2 focuses on data generation and method of the analysis. Section 3 discusses the empirical results and Section 4 concludes the study with policy recommendations.

2 | DATA AND METHODOLOGY

2.1 | Data

The data used in this study include: cases of COVID-19 infections in Nigeria extracted from the daily situation reports published by the Nigeria Center for Disease Control (NCDC) (2020) and recorded on the WHO (2020) portal; daily stock market returns (SMR) computed from the All Share Index (ASI) data published by the Nigerian Stock Exchange (NSE) (see Equation 1); daily inflation rate (INF) decomposed from the monthly inflation rate reported by the National Bureau of Statistics (NBS). Our sample covers the period between February 27, 2020 and April 30, 2020 (64 observations) due to data availability. The February 27, 2020 base is adopted because it was the date the index case of COVID-19 was announced in Nigeria, while April 30, 2020 is the date for latest available inflation data in Nigeria released by the NBS.

$$\text{SMR}_t = 100 \times \ln \left( \frac{\text{ASI}_t}{\text{ASI}_{t-1}} \right)$$

where SMR denotes stock market returns, ASI denotes all shares index, t denotes time period.

2.2 | Model specification and justification

Financial time series such as the stock market returns have three features that make modeling and analysis with it different from other normal time series. These include; volatility clustering, leptokurtic distribution, and leverage effect. For example, during a financial shock or crisis such as the one triggered by COVID-19, the leverage effect (or the asymmetric quality of financial time series data) is accentuated, thereby making it impossible to model volatility by normal means, rather, time-varying volatility models such as the ARCH or GARCH type models are used (Rastogi, 2014; Yousef, 2020). The approach of incorporating the time-varying nature of volatility using the ARCH
process was first suggested by Engle (1982). Further, the GARCH type models (GARCH, EGARCH, GJR-GARCH, and TGARCH) were developed by Bollerslev (1986) Glosten, Jagannathan, and Runkle (1993), Franses and Dick (1996) and Franses and Dick (2000) to overcome some of the limitations of ARCH models, such as overfitting and breach of the non-negativity constraint. To this effect, the GARCH family of models has become the standard method for modeling volatility in financial time series. This study, therefore, employs the standard GARCH—GARCH (1,1) and the GJR-GARCH model. The GARCH (1,1) model is asymmetric model, often used as the starting point for a volatility analysis. The GJR-GARCH on the other hand is an asymmetric model, believed to have the best fit among the competing family of volatility analysis. The GJR-GARCH model with inflation and COVID-19 in the variance equation is as follows:

$$\sigma_t^2 = \omega + (\gamma_1 + \delta \sigma_{t-1}) \sigma_{t-1}^2 + \lambda_1 \sigma_{t-1}^2 + \varphi_1 \text{INF}_t + \varphi_2 \text{COVID}_t$$  \hspace{1cm} (2)

where $\delta$ is the asymmetry or leverage effect parameter, and when $\delta > 0$, and significant; there is asymmetry or leverage effect, and negative shocks (bad news) will increase the volatility of the stock market returns more than good news. But if $\delta = 0$, the model collapses to GARCH (1,1) indicating that there is no leverage effect. The coefficients $\gamma_1$, $\lambda_1$ and represent the ARCH (recent news) and GARCH (old news) effects respectively. In addition, a higher value of $\lambda_1$ the GARCH coefficient implies that it will take a long time for shocks to conditional variance to dissipate or die off, suggesting that volatility is persistent. Also, if the sum of the ARCH and GARCH effects ($\gamma_1 + \lambda_1$) is close to 1, or relatively high, then a ‘shock’ at time $t$ will persist, which indicates a condition known as “long memory” or persistent of the conditional variance.

Furthermore, we apply the accounting innovation tests described in and Usman, Alola, and Sarkodie (2020) and Usman, Olanipekun, Ilorember, and Abu-Goodman (2020) which consist of impulse response functions (IRFs) and forecast error variance decomposition (FEVD) to analyze the response of both stock market returns and inflation to COVID-19 shocks to capture the short-run dynamics of the variables.

3 | RESULTS AND DISCUSSIONS

3.1 | Descriptive statistics

Table 1 presents information on descriptive statistics and correlation matrix for stock market returns (SMR), coronavirus (COVID-19), and inflation (INF). The results of the standard deviations in the descriptive statistics show that SMR and COVID-19 are the most volatile variables compared to INF because their values are farther away from the mean. The results also show that the highest number of COVID-19 infection cases recorded during the study period was 204. The results of the correlation matrix show a slightly strong and statistically significant correlation between SMR and INF and between INF and COVID-19, while the correlation between SMR and COVID-19 is weak but statistically significant. This finding excludes the risk of multicollinearity among the variables; hence, we conclude that there is no multicollinearity in the model. Figure 1 shows the time plot of the variables over the period of the study. Figure 1 further buttresses the volatile nature of SMR and COVID-19 over time while INF rises steadily over time. Further, Figure 2 illustrates the growth rate of COVID-19 and SMR and displays evidence of fluctuations in the variables over the study period. The initial wide negative fluctuation of SMR reflects investors’ bearish sentiments from February 27 on the announcement of the index case and the uncertainty over the impact of the regulatory and fiscal responses to the COVID-19 pandemic.

3.2 | Econometric analysis

Precursory to the application of the GARCH type models, we checked for unit root properties of the variables and establish that all the variables (SMR, INF, and COVID-19) are stationary at both l(0) and l (1) processes. Similarly, we tested for the presence or otherwise of heteroskedasticity in the stock market returns series using the autoregressive conditional heteroskedasticity (ARCH) or ARCH effect as a pre-requisite for using the volatility modeling approach. The result of the ARCH-LM test as evidence in Table 2 indicates that the stock market returns series is highly volatile since the null hypothesis of no ARCH effect is rejected at 5% level of significance. This finding, therefore, permits the application of the GARCH family of models which include the standard GARCH (GARCH 1,1) and the GJR-GARCH models for the purpose of this study.

Consequent to the results of the ARCH-LM test which warrants the application of the GARCH type models, we proceed to estimate the mean and conditional equations of the standard GARCH (1,1) model with INF and COVID-19 as the conditional variance regressors.

**TABLE 1 Descriptive statistics and correlation matrix**

|               | SMR  | INF  | COVID19 |
|---------------|------|------|---------|
| Mean          | -0.330605 | 12.29875 | 33.43478 |
| Std. Dev.     | 1.590511  | 0.047623 | 53.63235 |
| Maximum       | 2.971815  | 12.38405 | 204.0000 |
| Minimum       | -5.032939 | 12.22293 | 0.000000 |
| Correlation Matrix |
| SMR | 1
| INF | 0.514830 | 1 (0.0000) |
| COVID-19 | 0.228431 | 0.679791 | 1 (0.0000) |

Source: Authors’ computation.
The results are presented in Table 3. As expected, the results indicate that both the ARCH effect and the GARCH effect representing recent and past news respectively are statistically significant at a 5% level. This implies that both recent and past news have a significant impact on stock market returns volatility. In addition, the sum of the ARCH $\gamma$ (ARCH effect) and GARCH $\lambda$ (GARCH effect) effects is close to unity which suggests that the current shocks to the market returns triggered by COVID-19 may not die off speedily but rather persist for a while considering that the duration of the pandemic is unknown. That is to say that, the stock market volatility may continue even after the scourge of the dreaded coronavirus as the bearish sentiments are expected to last a little longer.

Further, the results in Table 3 indicate a positive and statistically significant impact of INF and COVID-19 on SMRs volatility. This suggests that both INF and COVID-19 have increased the impact of SMR volatility as investors scramble to adjust their portfolios to reflect the potential damage of the ‘virus’ to the global economy and asset prices.

Furthermore, it is observed that a shock to the stock market produces a disproportionate impact. That is, market downturns are
followed by higher volatility, a scenario known as leverage effect.
Leverage or asymmetric effect exists if bad news produces higher
stock market volatility than good news. Since GARCH (1,1) model do
not account for the leverage effect, it is imperative to consider a
model that accounts for the leverage effect in the stock market vola-
tility, hence we estimate the GJR-GARCH (1,1) model to ascertain the
leverage effect and the results are presented in Table 4. Just like the
results of the GARCH (1,1) model, the asymmetric GARCH results
show that INF and COVID-19 significantly fuel stocks market returns
volatility, and there is long-memory as to the effect of the shock.
Regarding the leverage effect which is the basic reason for applying
the GJR-GARCH test, the results show that the coefficient of the
leverage effect is positive and statistically significant at a 5% level,
which suggests the presence of leverage or an asymmetric effect.
That is, a negative shock such as the current COVID-19 pandemic and
the associated INF increases SMR volatility more than a positive
shock. This further shows that COVID-19 distorts the positive rela-
tionship between stock market returns and inflation as postulated by
the Fisher hypothesis but rather aligns with the Fama proxy hypothe-
sis. This finding is consistent with the findings of Yousef (2020) which
establishes that COVID-19 increases SMR volatility across the three
stocks markets (S&P500, Dow Jones, and Nasdaq) in the United
States of America.

Table 5 presents the estimates of the VAR model. Even though
the interpretations of the estimates of VAR may not be meaningful
because of the theoretical nature of this model, the study relies on
the forecast error variance decompositions and impulse response
functions for insightful interpretation as suggested by Sim (1980) in

### Table 2: Results of unit root and ARCH effect tests

| Variable | Test Statistic | 5% CV | ARCH Effect Test Statistic | 5% CV | LM
|----------|---------------|-------|---------------------------|-------|-------
| SMR      | −3.7863**     | −2.9350 | −5.6012**                       | −2.9369 | 4.7364** |
| INF      | −4.8977**     | −2.9350 | −6.2890**                     | −2.9369 | -      |
| COVID_19 | −2.9764***    | −2.9350 | −4.7785**                     | −2.9369 | -      |

Note: ** indicates significance at 5% level of significance. CV implies critical values and LM stands for Lagrange multiplier.

### Table 3: Results of GARCH (1,1) with inflation and COVID-19

| Coefficient       | p value |
|-------------------|---------|
| μ                 | .5909*** .0073 |
| ω                 | .9621** .0717 |
| γ (ARCH effect)   | .2486** .0349 |
| δ (GARCH effect)  | .6213** .0294 |
| γ + δ (Persistence)| .8699 |
| INF               | .8712** .0286 |
| COVID-19          | .4693*** .0034 |

Note: *** and ** indicate 1 and 5% level of significance.

### Table 4: Results of GJR-GARCH with inflation and COVID-19

| Coefficient       | p value |
|-------------------|---------|
| μ                 | .3027** .0390 |
| ω                 | 1.0372** .0362 |
| δ (Leverage effect) | .5199* .0480 |
| γ (ARCH effect)   | .2671** .0244 |
| δ (GARCH effect)  | .7129*** .0001 |
| γ + δ (Persistence)| .9799 |
| INF               | .7185** .022 |
| COVID-19          | .1833** .005 |

Note: ***, ** and * indicate 1, 5 and 10% level of significance.

### Table 5: Estimates of the VAR model

| Variables       | SMR          | INF          | COVID-19     |
|-----------------|--------------|--------------|--------------|
| SMR             | 0.0977       | 0.0003       | 0.0134       |
| INF             | 0.1018       | 0.0002       | 0.1018       |
| COVID-19        | 4.5782***    | 1.0170       | 4.4582***    |
| Constant        | −1.7117      | −0.1909      | −1.4862      |
| R squared       | 0.2669       | 0.9983       | 0.8684       |
| Adj. R squared  | 0.1479       | 0.9981       | 0.8470       |
| F statistic     | 2.2448       | 3.6993       | 40.680       |
| Akaike AIC      | 3.7494       | −9.4240      | 9.0969       |
| Schwarz SC      | 4.0333       | −9.1402      | 9.3807       |

Note: *** denotes significance 1% while standard errors are presented in parentheses.
Balcilar, Roubaud, Usman, and Wohar (2020). Based on the estimated results, the study finds positive and negative effects among the variables with evidence of statistical significance. The results of the impulse response functions (IRFs) display in Figure 3 reveal the responses of SMR to own innovation shock, INF shock, and COVID-19 shock using a 95% confidence interval. These responses are estimated based on the period of 10 days, which are robust enough to assess the full effects of INF and COVID-19 on the Nigeria SMR. The results show that SMR responds positively to an increase in its own innovation shock and the response is statistically significant up to the second horizon; after which it becomes insignificant. The study further observes that the response converges to its equilibrium value after the seven horizons. More so, the response of SMR to INF shock is negative and insignificant. The response becomes ineffective and neutral after five horizons. Finally, the study finds that SMR responds to COVID-19 shock negatively over the study period. This response is ineffective during periods of first and second horizons; after which it becomes effective. The negative impact of COVID-19 on the SMR justifies the fact that it is difficult to predict with certainty the magnitude of its impact on the Nigerian economy as market analyst expects the pandemic to last longer as already established in both the GARCH (1,1) and GJR-GARCH results.

Further, the estimated results of the forecast error variance decomposition (FEVD) are displayed in Table 6. It shows that the contribution of SMR to own shock is 90.88%. This is followed by the contributions from COVID-19 of 7.27%, and INF of 1.85% to the SMR forecast error variance decomposition. For INF, we find that the contribution to own shock is the largest with 92.32%, followed by

| Period | S.E. | SMR | INF | COVID19 |
|--------|-----|-----|-----|---------|
| 1      | 1.467142 | 0.000000 | 0.000000 | 1.000000 |
| 2      | 1.488827 | 98.48509 | 1.453872 | 0.061035 |
| 3      | 1.544952 | 97.22563 | 1.697149 | 1.077224 |
| 4      | 1.554762 | 96.74464 | 1.911476 | 1.343871 |
| 5      | 1.565955 | 95.68395 | 1.939331 | 2.376720 |
| 6      | 1.573773 | 94.90944 | 1.928798 | 3.161764 |
| 7      | 1.582349 | 93.91087 | 1.911692 | 4.177437 |
| 8      | 1.590923 | 92.93157 | 1.891146 | 5.177285 |
| 9      | 1.599759 | 91.91422 | 1.870342 | 6.215438 |
| 10     | 1.608886 | 90.88101 | 1.849205 | 7.269787 |

Variance Decomposition of INF

| Period | S.E. | SMR | INF | COVID19 |
|--------|-----|-----|-----|---------|
| 1      | 0.002023 | 3.398823 | 96.60118 | 0.000000 |
| 2      | 0.002631 | 2.090842 | 97.73218 | 0.176974 |
| 3      | 0.003138 | 2.236676 | 97.40600 | 0.357329 |
| 4      | 0.003617 | 1.881154 | 97.23571 | 0.883140 |
| 5      | 0.004025 | 1.820253 | 96.78938 | 1.390365 |
| 6      | 0.004424 | 1.751342 | 96.13105 | 2.117608 |
| 7      | 0.004792 | 1.699011 | 95.37614 | 2.924850 |
| 8      | 0.005150 | 1.661148 | 94.46663 | 3.872221 |
| 9      | 0.005495 | 1.618547 | 93.45207 | 4.929385 |
| 10     | 0.005832 | 1.580026 | 92.32030 | 6.099678 |

Variance Decomposition of COVID19

| Period | S.E. | SMR | INF | COVID19 |
|--------|-----|-----|-----|---------|
| 1      | 21.26473 | 0.035684 | 4.788704 | 95.17561 |
| 2      | 25.18841 | 0.602702 | 10.02927 | 89.36803 |
| 3      | 29.75137 | 0.522268 | 7.198011 | 92.27972 |
| 4      | 33.19393 | 0.770548 | 6.831580 | 92.39787 |
| 5      | 36.42628 | 0.641762 | 6.203465 | 93.15477 |
| 6      | 39.42475 | 0.655470 | 5.907505 | 93.43702 |
| 7      | 42.21288 | 0.608397 | 5.803977 | 93.58763 |
| 8      | 44.87181 | 0.596150 | 5.755608 | 93.64824 |
| 9      | 47.39188 | 0.582533 | 5.827666 | 93.58980 |
| 10     | 49.80896 | 0.575407 | 5.941861 | 93.48273 |

Cholesky ordering: SMR INF COVID19

![Figure 3: Impulse response function to inflation and COVID-19 shocks](image)
COVID-19 with 6.10% and lastly SMR with 1.58%. Finally, we observe that the contribution of COVID-19 to own shock is 93.48% (which is the largest), followed by contributions from INF and SMR shocks of 5.94 and 0.58% respectively.

4 | CONCLUSION AND RECOMMENDATIONS

The study examines the nexus between stock market returns and inflation rate in Nigeria by accounting for the effect of the novel coronavirus (COVID-19). The study sets out to determine whether stock market returns hedge against inflation as espoused by the Fisher (1930) hypothesis or as the reverse as established by the Fama (1981) proxy hypothesis. The study deploys the GARCH (1,1) and GJR-GARCH as well as the accounting innovation tests on daily data from February 27, 2020 to April 30, 2020 covering the date the COVID-19 index case was announced in Nigeria to the latest date (month) of inflation rate data. The results of the GARCH (1,1) and the GJR-GARCH models divulge that both inflation rate and COVID-19 cases increase stock market returns volatility. The findings of the GJR-GARCH model as reflected in the leverage effect further establishes that the negative shock accentuated by the rising number of COVID-19 infection cases increases volatility and distorts the positive relationship between inflation and stock market returns, implying a negation of the Fisher hypothesis and aligning with the Fama Proxy Hypothesis. More so, the results of the impulse response and the forecast error variance decomposition indicates that stock market returns responds negatively to COVID-19 shock over the study period. The effectiveness of the response is however progressive with the passage of time, which suggests that the negative effects of COVID-19 on the market returns and its disruption to the stock market returns—inflation relationship is one that would not die away rapidly considering that the duration of the pandemic is unknown.

The study therefore recommends the need for the fiscal and regulatory authorities to strengthen collaboration on policy measures that will stabilize and reposition the stock market to continue in its critical role of galvanizing funds for investment and stimulating growth. Some of these measures may include: activating business continuity plans and health precautionary measures to ensure that trading activities continued seamless remotely without exposing staff and traders to threat of the pandemic; issuing guidance to capital market operators and other stakeholders on the operations of the stock market activities during the pandemic, sustained monitoring of the capital market activities to ensure that high ethical standards are maintained; ensuring that relevant information continues to flow into the market space through reliable and transparent medium such as official websites, electronic and social media platforms to enable investors make informed investment decisions. In addition, the study recommends expansion of intervention measures to cover sectors such as Transportation (Aviation Sector), Hospitality and Tourism, Construction, Entertainment given that the trading activities of companies listed in these sectors on the Nigerian Stock Exchange contributes to its daily market performance (All Share Index) and Stock Market Returns.

ORCID
Gylych Jelilov  https://orcid.org/0000-0002-7428-7361
Paul Terhemba Iorember  https://orcid.org/0000-0002-3092-2378
Ojunughwa Usman  https://orcid.org/0000-0002-6459-9898
Paul M. Yu  https://orcid.org/0000-0003-3542-0892

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**AUTHOR BIOGRAPHIES**

**Gyllych Jellov** is an Associate Professor of Economics at Nile University of Nigeria, Abuja. He holds a PhD in Economics and conducts research in Macro and Micro Economics, Development Economics, Public Policy, Energy economics, Renewable Energy, Environmental Economics as well as Macroeconomic modeling. He is well published in reputable scientific journals including African Development Review, Journal of Public Affairs, Energy Journal amongst others.

**Paul Terhemba Iorember** is a PhD Scholar in Economics from the University of Jos, Nigeria. He does research in Macroeconomics, Energy finance, Renewable Economics, International Finance, Applied Econometrics, as well as Macroeconomic modeling (Macroeconometric modeling and Computable General Equilibrium Modeling). He is well published in reputable scientific journals including African Development Review; Environmental Science and Pollution Research; Journal of Public Affairs; International Journal of Strategic Energy and Environmental Planning; American Journal of Economics, African Journal of Economic Review amongst others.

**Ojonugwa Usman** holds a PhD in Economics from Eastern Mediterranean University, North Cyprus. He researches in International Economics and Finance, Open Economy Macroeconomics, Energy and Environmental Economics, Corporate Governance and Applied Econometrics. He has published scientific research papers in reputable journals including Current Issues in Tourism, Journal of Environmental Economics and Social Sciences, Science of the total environment, African Development Review, Romanian Journal of Economic Forecasting, The World Economy, Journal of Public Affairs, amongst others.
Paul M. Yua holds a PhD in Banking and Finance from Nnamdi Azikiwe University Akwa, Nigeria. He researches in Banking and Finance, Capital and Money markets, Risk Management, Corporate Finance, Supervision and Regulation. He is well published in reputable Finance and Economic Journals.

**How to cite this article:** Jelilov G, Iorember PT, Usman O, Yua PM. Testing the nexus between stock market returns and inflation in Nigeria: Does the effect of COVID-19 pandemic matter? *J Public Affairs*. 2020;20:e2289. [https://doi.org/10.1002/pa.2289](https://doi.org/10.1002/pa.2289)