Insurability of HIV-positive people treated with antiretroviral therapy in Europe: collaborative analysis of HIV cohort studies

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Objective: To increase equitable access to life insurance for HIV-positive individuals by identifying subgroups with lower relative mortality.

Design: Collaborative analysis of cohort studies.

Methods: We estimated relative mortality from 6 months after starting antiretroviral therapy (ART), compared with the insured population in each country, among adult patients from European cohorts participating in the ART Cohort Collaboration (ART-CC) who were not infected via injection drug use, had not tested positive for hepatitis C, and started triple ART between 1996–2008. We used Poisson models for mortality, with the expected number of deaths according to age, sex and country specified as offset.

Results: There were 1236 deaths recorded among 34 680 patients followed for 174 906 person-years. Relative mortality was lower in patients with higher CD4 cell count and lower HIV-1 RNA 6 months after starting ART, without prior AIDS, who were older, and who started ART after 2000. Compared with insured HIV-negative lives, estimated relative mortality of patients aged 20–39 from France, Italy, United Kingdom, Spain and Switzerland, who started ART after 2000 had 6-month CD4 cell count at least 350 cells/\textmu{}l and HIV-1 RNA less than$10^4$ copies/ml and without prior AIDS was 459%. The proportion of exposure time with relative mortality below 300, 400, 500 and 600% was 28, 43, 61 and 64%, respectively, suggesting that more than 50% of patients (those with lower relative mortality) could be insurable.
Conclusion: The continuing long-term effectiveness of ART implies that life insurance with sufficiently long duration to cover a mortgage is feasible for many HIV-positive people successfully treated with ART for more than 6 months.

Methods

The ART-CC, described in detail elsewhere [9,10], is an international collaboration of HIV cohort studies that combines data on HIV-positive individuals who were antiretroviral-naive when they started ART. Here, eligible patients were participants in European cohorts, aged at least 18 years, had presumed mode of transmission that was not via IDU, and initiated ART between 1996 and 2008. All had CD4 cell count and HIV-1 RNA measured in the 3 months before ART initiation, and between 3 and 9 months after initiation. The included cohorts, from six countries, were: the AIDS Therapy Evaluation Project Netherlands (ATHENA); the Agence Nationale de Recherches sur le SIDA et les hépatites virales (ANRS) CO4 French Hospital Database on HIV (FHDH) and ANRS CO3 Aquitaine Cohort, France; Italian Cohort of Antiretroviral-Naive Patients (ICONA); Royal Free Hospital Cohort, London United Kingdom; Swiss HIV Cohort Study (SHCS); the Cohorte de la Red de Investigacion en Sida (CoRIS), the Proyecto para la Informatización del Seguimiento Clínico-epidemiológico de la Infección por HIV y SIDA (PISCIS) Cohort, VACH, Spain; and the multicenter study, EuroSIDA (restricted to patients from the selected six countries). Data on follow-up for mortality were available to 31 December 2009. Patients not known to have died were considered lost to follow-up 3 months after their last CD4 cell count measurement, except when this was within 6 months of study end. Institutional review boards from each cohort approved analysis of routinely collected data.

Expected numbers of deaths according to age, sex and country were calculated based on so-called ‘standard’ insured lives (those healthy lives accepted for life insurance using standard premium rates without a health loading). These were based on published insurance tables where available (France, Netherlands, United Kingdom), or on adjustments to population mortality tables (Italy, Spain, Switzerland) commonly used in those markets by pricing actuaries. The specific percentages used for each market (country) were obtained from interviews with Swiss Re (a reinsurer with headquarters in Zurich, Switzerland) regional pricing actuaries and were based on their knowledge of pricing levels in those markets. Such multiplicative adjustments reflected typical industry pricing levels for individual life insurance in 2004, taking into account the risk assessment (underwriting) practices in 2011 and accounts for insured lives having lower mortality than general population lives. Insurance mortality trends are a combination of the underlying health and accident trends in the insured population and the trends in (or changes in practice of) how underwriters decide which lives belong to the pool of insured persons with standard risk. For France we used 50% of mortality rates in actuarial tables TF00-02 (women) and 40% of TH00-02 (men), respectively (www.institutesdesactuaires.com/docs/2007017232113_NOTICETHTF0002.pdf). For the Netherlands we used 55% of the respective GBM/V00-5 tables (www.ag-ai.nl/download/10351-Prognostetafel+2005+-+2050.pdf). For the UK we used 70% of TF00 and 60% of TM00.
respectively (www.actuaries.org.uk/research-and-resources/pages/00-series-mortality-tables-assured-lives-annuitants-and-pensioners). For Italy we used 55% of the respective 2002 population mortality tables for each sex published by the Italian National Institute of Statistics (table SI 2002 www.istat.it). For Spain we used 45% of the 2006 population mortality and for Switzerland we used 80% of the 2006 population mortality (Human Mortality Database www.mortality.org).

The United Kingdom and Netherlands tables provided rates by age and duration at policy inception: the effect of duration generally decreases with increasing time since the underwriting date and the rates after which there is no substantial change with increasing duration are termed ‘ultimate’. The lives in the HIV study were not screened through the normal insurance processes. It is, thus, consistent to compare them with insurance rates where the effects of such screening have worn off, which correspond to the ‘ultimate rates’ in the comparator insured population. Therefore, we used the ultimate rates, which depend only on age, from these tables.

We graphed assumed mortality rates for insured lives (using the above-mentioned tables) by age, sex and country, and compared mortality rates for insured lives and the general population by age and sex. The general population reference was based on the human mortality database (www.mortality.org) for 2004, and the rates for each country weighted according to the person-years of observation in the analysis dataset.

Data analyses used generalized linear models (GLM) assuming a Poisson error distribution with log link function [11] and with expected numbers of deaths based on the insured population (person-years at risk multiplied by the rate in the corresponding insured population) according to age, sex and country specified as an offset [12]. As we wished to model duration of ART explicitly, we did not allow for it in the offset. The variables considered for inclusion were country, sex, transmission risk group (MSM, heterosexual contact, blood and other), year of initiation of ART, age at start of ART, AIDS diagnosis prior to ART and CD4 cell count and HIV-1 RNA at start of ART and 6 months later (see Table 1 for categorizations). The 6-month CD4 cell count and HIV-1 RNA used in the analysis were those measured closest to 6 months after start of ART and within a window of 3–9 months. We additionally considered for inclusion in the GLM three variables that varied over time: calendar time (1996–2000, 2001–2008), current age (20–29, 30–39, 40–49, 50–59, ≥60 years) and duration of ART since the 6-month measurement (1–3, 4–6, 7–9, ≥10 years). Note that duration of ART equal to 1–3 corresponds to the first 3 years after the 6-month measurement, and therefore is follow-up time from 0.5 to 3.49 years after starting ART.

Starting with a model including all variables, we removed variables to find a model containing only variables with clear associations (P < 0.05) with mortality and where the residuals were not overly overdispersed or underdispersed. The dispersion parameter measures both underdispersion and overdispersion and should be close to one for a well fitting Poisson model. We combined categories where this improved model fit, and tested two-way, but not higher-order, interactions. We chose the final model based on dispersion parameter within 5% of 1, lower Akaike Information Criterion (AIC) [13] and normality of the distribution of Anscombe residuals [11]. In sensitivity analyses we used P value less than 0.1 as the threshold for variable selection and considered models with a negative binomial distribution.

The final GLM was used to estimate the relative mortality of the group with reference values of all included variables, compared with insured lives within the same age and country group. The relative mortality of other groups was derived by multiplying baseline relative mortality by mortality rate ratios from the GLM. Potential insurability is determined by the actual/expected claims ratio (relative mortality multiplied by 100%). A claims ratio of 100% represents no excess mortality, whereas 500%, or equivalently relative mortality 5, represents 400% excess mortality. Bounds for insurability are not fixed, and therefore we illustrated the extent of excess mortality by drawing plots of actual/expected claims ratios with contours at 100, 250, 500, 750, 1000, 2000 and 3000% relative mortality, according to 6-month CD4 cell count, ART duration, age and calendar period of ART initiation, among people with lower risk values of other variables. Stata® version 12 and Microsoft Excel 2007 were used to perform analyses.

**Results**

Among 34 680 patients followed for 174 906 person-years there were 1236 deaths (overall mortality rate 0.71 [95% confidence interval (CI) 0.67–0.75] per 100 person-years). The majority of the data were from France and the Netherlands (Table 1) and 70% of patients were men. Crude mortality rates increased with age and were higher for men and those with an AIDS diagnosis prior to baseline. The association of lower CD4 cell count and higher HIV-1 RNA with mortality was greater for 6-month measurements than those at initiation of ART. Crude mortality rates varied between countries and were lower for the United Kingdom and Italian cohorts, which may reflect true lower mortality or might be due to sampling variability or incomplete death ascertainment.

Figure 1 (upper panel) shows mortality rates in the insured population according to sex, age and country. Differences in mortality rates between countries were
relatively small, UK men and women had the highest mortality rates at most ages, whereas Spanish women had the lowest mortality. Figure 1 (lower panel) shows the lower mortality in the insured population compared with the whole population, in both men and women.

The final model did not include sex: the model offset effectively adjusted for sex as mortality rates were lower in women in the insured population. CD4 cell count and HIV-1 RNA at initiation of ART were not sufficiently prognostic for inclusion in the final model, after adjustment for 6-month values of those variables. Age at initiation was not prognostic after adjustment for current age. There was little difference between mortality from 7–9 and at least 10 years duration of ART and so these categories were combined. Similarly, categories for starting ART in 2001–2004 and 2005–2008 were combined, and HIV-1 RNA was dichotomized at 10^4 copies/ml. The dispersion parameter was lowest in models that included all countries separately, and therefore groupings of countries were examined. Comparing the two countries contributing most person-years, Dutch lives experienced higher adjusted relative mortality than French lives, partly because of the lower standard mortality of the Dutch insured HIV-negative population. We found little evidence that relative mortality in the other included countries differed from that of either France or the Netherlands, but that the best-fitting model was obtained by grouping other countries with France. After choosing prognostic variables and their groupings and restricting to models with dispersion parameter within 5% of 1, two models with similar AIC and Anscombe residual distribution remained. We chose the model that did not include transmission risk category, as differential pricing of insurance based on slightly higher mortality rates in individuals with transmission via blood transfusion would be difficult to justify. The final model did not include any interactions. Sensitivity analyses indicated that variables selected for inclusion in the final model were the same when the P value threshold was

### Table 1. Characteristics of 34 680 HIV-positive patients included in the analyses.

| Characteristic                        | Category (Character) | Number of patients (%) | No. of deaths (%) | Crude mortality rate per 100 person-years (95% CI) |
|--------------------------------------|----------------------|------------------------|------------------|-----------------------------------------------|
| Total                                |                      | 34 680 (100)           | 1236 (100)       | 0.71 (0.67–0.75)                              |
| Country                              |                      |                        |                  |                                               |
| France                               |                      | 19 503 (56)            | 684 (55)         | 0.71 (0.66–0.76)                              |
| Italy                                |                      | 1683 (5)               | 38 (3)           | 0.43 (0.29–0.58)                              |
| Netherlands                          |                      | 5718 (16)              | 250 (20)         | 0.78 (0.69–0.89)                              |
| Spain                                |                      | 3379 (10)              | 85 (7)           | 0.61 (0.48–0.75)                              |
| Switzerland                          |                      | 3209 (9)               | 147 (12)         | 0.82 (0.69–0.96)                              |
| UK                                   |                      | 1188 (3)               | 32 (3)           | 0.53 (0.36–0.74)                              |
| Sex                                  | Female               | 10 504 (30)            | 216 (17)         | 0.44 (0.38–0.50)                              |
|                                       | Male                 | 24 176 (70)            | 1020 (83)        | 0.81 (0.76–0.86)                              |
| Age at initiation of ART (years)     | 15–29                | 6891 (20)              | 118 (10)         | 0.34 (0.28–0.41)                              |
|                                       | ≥30–49               | 22 437 (65)            | 685 (55)         | 0.59 (0.55–0.64)                              |
|                                       | ≥50                  | 5352 (15)              | 433 (35)         | 1.71 (1.54–1.87)                              |
| Year of ART initiation                | 1996–2000            | 13 005 (38)            | 765 (62)         | 0.76 (0.71–0.81)                              |
|                                       | 2001–2004            | 12 807 (37)            | 375 (30)         | 0.64 (0.57–0.71)                              |
|                                       | 2005–2008            | 8868 (26)              | 96 (8)           | 0.62 (0.50–0.75)                              |
| Transmission risk group               | MSM                  | 13 845 (40)            | 497 (40)         | 0.66 (0.60–0.72)                              |
|                                       | Heterosexual         | 17 948 (52)            | 580 (47)         | 0.67 (0.62–0.73)                              |
|                                       | Blood                | 335 (1)                | 24 (2)           | 1.29 (0.83–1.93)                              |
| Baseline CD4 cell count (cells/μl)   | >350                 | 8200 (24)              | 186 (15)         | 0.40 (0.34–0.46)                              |
|                                       | 200 to 349           | 10 669 (31)            | 266 (22)         | 0.54 (0.47–0.60)                              |
|                                       | 100 to 199           | 7020 (20)              | 263 (21)         | 0.79 (0.69–0.89)                              |
|                                       | 50 to 99             | 3404 (10)              | 175 (14)         | 1.00 (0.85–1.16)                              |
| Six-month CD4 cell count (cells/μl)  | >350                 | 17 001 (49)            | 349 (28)         | 0.40 (0.36–0.44)                              |
|                                       | 200 to 349           | 9425 (27)              | 312 (25)         | 0.68 (0.61–0.76)                              |
|                                       | 100 to 199           | 5689 (16)              | 275 (22)         | 0.95 (0.83–1.06)                              |
|                                       | 50 to 99             | 1788 (5)               | 147 (12)         | 1.57 (1.32–1.84)                              |
| Baseline HIV-1 RNA (copies/ml)       | <50                  | 777 (2)                | 153 (12)         | 4.61 (3.88–5.37)                              |
|                                       | ≥500 to 1 x 10^4     | 5033 (15)              | 448 (18)         | 0.58 (0.49–0.68)                              |
|                                       | ≥1 x 10^5 to 1 x 10^6| 12 265 (35)            | 376 (30)         | 0.59 (0.53–0.66)                              |
| Six-month HIV-1 RNA (copies/ml)      | <500                 | 24 725 (71)            | 685 (55)         | 0.60 (0.56–0.65)                              |
|                                       | ≥500 to 1 x 10^4     | 7298 (21)              | 313 (25)         | 0.65 (0.58–0.72)                              |
|                                       | ≥1 x 10^5 to 1 x 10^6| 1788 (5)               | 117 (9)          | 1.33 (1.09–1.58)                              |
| Clinical CDC stage at baseline       | A/B                  | 27 067 (78)            | 749 (61)         | 0.55 (0.51–0.59)                              |
|                                       | C                    | 7613 (22)              | 487 (39)         | 1.25 (1.14–1.36)                              |

ART, antiretroviral therapy; CDC, Centers for Disease Control.
Mortality rate ratios for the variables included in the final model are displayed in Table 2, which shows that CD4 cell count and viral load at 6 months were the most strongly prognostic variables. Mortality rate ratios decreased with age \((P < 0.005)\) because increases in mortality with age in HIV-positive people were less marked than in comparator populations. Mortality rate ratios decreased with ART duration and were higher for those who initiated ART before 2001 \((P < 0.005)\) or who had an AIDS diagnosis prior to starting ART \((P < 0.005)\).

The baseline group for comparison of mortality ratios was patients aged 20–39 from France, Italy, United Kingdom, Spain and Switzerland, who started ART after 2000, had 6-month CD4 cell count at least 350 cells/µl, 6-month HIV-1 RNA < 10^4 copies/ml and no AIDS before ART initiation, during the first 3 years of ART. The relative mortality of this group compared with insured HIV-negative lives in the same age group was 459\% (95% CI 391–539). The relative mortality of other groups is derived by multiplying by the corresponding mortality rate ratios in Table 2. For example, patients aged 40–49 years with 6-month CD4 cell count 200–349 cells/µl, who started ART before 2001 and otherwise the same characteristics as the baseline group had relative mortality 459\% × 0.53 × 1.44 × 1.33 = 466\% during the

Fig. 1. Mortality rates according to age and country of insurance population (upper panels) and weighted European mortality comparing insurance population and general population (lower panels).
first 3 years, \(466\% \times 0.69 = 321\%\) during years 4–6 and \(466\% \times 0.51 = 238\%\) after 7 or more years on ART.

Figure 2 shows contours of relative mortality according to CD4 cell count and duration of treatment, separately according to age group and calendar period of starting ART, among patients with lower risk values of other variables. Supplementary Table 1, http://links.lww.com/QAD/A325 shows percentage relative mortality according to risk factors for France, Italy, Spain, Switzerland, and United Kingdom (illustrated in Fig. 2) and supplementary Table 2, http://links.lww.com/QAD/A325 shows relative mortality for the Netherlands. In our study population (which excluded IDU or those who were hepatitis C positive), the proportion of exposure time with relative mortality below 300, 400, 500 and 600% was 28, 43, 61 and 64%, respectively.

**Discussion**

We estimated the relative mortality of HIV-positive lives, compared with insured HIV-negative lives in the same population. Our results apply to HIV-positive individuals who have survived at least 6 months after starting ART in six European countries. Relative mortality was lower for those who started ART after the year 2000 without a prior AIDS diagnosis and who had 6-month CD4 cell count greater than 350 cells/\(\mu\)l and 6-month HIV-1 RNA less than \(10^4\) copies/ml. Relative mortality compared with insured HIV-negative lives declined with increasing duration of ART, and decreased with age despite increases in mortality rates with age, a phenomenon that has been observed in other studies of HIV populations [14–16]. Because relative mortality varies with age and with time on ART, no single result applies to any particular insurance policy, our results must be used as a

![Fig. 2. Percentage relative mortality compared with HIV-negative insured lives (actual/expected claims ratio) of HIV-positive people according to CD4 cell count 6 months after start of ART, duration of ART, age group and calendar period of initiation of ART, among individuals without an AIDS diagnosis and with 6-month HIV-1 RNA less than 1000 copies/ml and in countries other than the Netherlands. Typical bounds for insurability range between 400 and 500%.](image_url)
set. Actuarial methods for converting varying mortality loadings so as to obtain a level of extra premium are beyond the scope of this article. However, our results imply that more than 50% of patients — those with lower relative mortality — in an HIV-positive population with similar risk profile to that analysed in this study could be insurable, based on conventional limits to relative mortality of 500%.

Our study was based on a large HIV-positive study population, we analysed data from over 30,000 patients, with long follow-up during which more than 1000 deaths were recorded. We controlled for underwriting effects by using ‘ultimate rates’, which factor out selection bias (see above for further details). Data were only available to estimate death rates up to 10 years after starting ART, and therefore estimated excess mortality will have to be extrapolated beyond the follow-up time that is currently observable. However, most HIV-related mortality occurs during the first few years of ART [17], and therefore excess mortality rates beyond 7 years duration of ART are likely to be stable. Our estimates are likely to be conservative for patients starting ART today on better tolerated and more effective drug regimens than those for patients included in analyses. A limitation of our study is that there was only limited follow-up of older individuals: 6572 person-years (3.8% of the total follow-up time) among persons aged more than 60 years. Therefore, death rates above age 60 were imprecisely estimated. However, demand for life insurance is mainly for lives under age 60 years. We excluded patients with presumed transmission via IDU, as these are unlikely to be granted insurance even if HIV negative. We did not have data on socioeconomic status. However, bias should be small because we used socioeconomic neutral mortality rates. We did not have data on smoking, this could lead to a bias if, for example, HIV-positive people were on average more likely to smoke than other insured people. The effect of this is likely to be relatively small, for example, a doubling of the proportion of smokers would result in an approximately 25% increase in mortality.

Our dataset included national databases on HIV-positive people from France, the Netherlands and Switzerland; these are likely to be representative of all patients in care in these countries. Data from Spain were contributed by many treatment centres from several regions, whereas those from Italy and Germany were from a limited number of clinics. All clinics provide comprehensive care to HIV-positive people in their locality and are not specialist referral centres. The majority of the data analysed in this study were from France and the Netherlands. We found relative mortality to be somewhat higher in the Netherlands, partly because of relatively lower mortality in Dutch insured HIV-negative lives. We found little evidence that relative mortality in other countries differed from that in France. We found little evidence that effects of variables included in the final model differed between Dutch lives and those from other countries. Our results may apply to other high-income countries with good access to care but it would be desirable to conduct further analyses of data from European countries not represented here, as insurance policies vary by country. In the USA, part of the HIV epidemic is in socially marginalized populations [18, 19], and access to free healthcare is limited, for example, only those with an AIDS diagnosis or with CD4 cell count less than 200 cells/µl are eligible for Medicaid, although provision will be expanded when the Affordable Care Act takes full effect in 2014 [20]. It is, therefore, unclear whether a similar proportion of HIV-positive lives is potentially insurable, and a study based on representative cohorts of HIV-positive people in the USA is highly desirable. Similarly, there is an urgent need for a study of insurability in South Africa, wherein the prevalence of HIV is many times higher than in Europe and there is also a high market penetration of life insurance. Studies are also lacking in Japan, India, China and Brazil — countries with potentially large insurance markets for HIV-positive lives.

People newly diagnosed with HIV can be expected to survive longer than those recruited to cohorts between 1996 and 2010, therefore, studies such as ours necessarily provide trailing indicators of mortality rates. Attempts have been made to allow for this problem in a computer simulation model: estimated life expectancy of MSM aged 30 years who became HIV positive in 2010 with no hepatitis C infection and high CD4 cell count at start of ART in the United Kingdom was 75 years (95% CI 68–77), a loss of only 7 years compared with the male UK population [21], similar to the average loss of life because of smoking.

Theoretically, insurability depends on the financial resources of the person buying insurance as insurance providers discriminate levels of risk, and price products accordingly. However, to prevent fraud and antiselection, insurers commonly apply their own insurability criteria which include limits on insurance ratings. Practices vary with a limit on the percentage extra mortality being the commonest. We have assumed that ratings up to an excess mortality of 400% (relative mortality 500%) are insurable. Insurance ratings beyond these bounds have generally not been accepted by the insurance buying public.

The Dutch Association of Insurers recommended in 2005 that short-term insurance should be made available to successfully treated patients [22] and in 2009 that it should be extended to patients who do not yet need ART [23]. Nevertheless, HIV-positive individuals have routinely been turned down for life insurance, despite advances in ART and evidence of its effectiveness [24]. Our study should allow insurance providers to make a fair assessment of risk in order to properly underwrite insurance for HIV-positive people, for longer terms. Assuming that relative mortality does not increase as time on ART increases.
beyond the maximum analysed here, our study provides evidence that could allow life insurance up to 20 years term to be offered to lower risk HIV-positive individuals, at affordable premiums. Whole of life insurance at guaranteed rates may become feasible when data on mortality with longer duration of ART become available.

Remaining underwriting challenges include assessment of likely future adherence to ART, the quality of medical care, presence of comorbidities and drug interactions. There are also pricing challenges due to extrapolation of mortality rates beyond the 10 years’ follow-up analysed here, and because HIV-positive individuals may apply for insurance at any time prior to or during treatment, but our estimates are based on CD4 cell count and HIV-1 RNA 6 months after ART initiation. Information on CD4 trajectories and viral suppression can be used to map from duration of ART to duration since taking out insurance [25]. Most chronic diseases result in deteriorating health with time on medication, but HIV-positive individuals on ART experience an increasing relative benefit of treatment as their CD4 cell counts increase towards normal levels, which may take more than 5 years [14,26]. Therefore, the lives of people with HIV tend to become more insurable with increasing duration of successful ART. Insurance companies set bounds for insurability for standard lives, including high age at application (which vary considerably between insurers and markets but could be as low as 60 years) as well as restrictions based on prior or current recreational drug use, which would also apply to HIV-positive lives. Our model suggests poorer outcomes for those with AIDS before start of ART. Insurers will have to interpret this with caution as different AIDS-defining conditions have substantially different implications for subsequent mortality [27].

Many people living with HIV and starting ART with currently recommended drugs and in accordance with treatment guidelines will live a near-normal life span [21]. Lack of insurance products is no longer justified, as the excess mortality of those with HIV is comparable to many other groups with morbidities that are insured, such as diabetics or cancer survivors. Our study provides data that will allow the insurance market to open up to people living with HIV. Life insurance could now be extended to 20 years, which would be particularly useful for mortgage cover. We aim to communicate our results directly to insurance providers in order that they can amend their policies, with consequent improvements in the quality of life for HIV-positive people.

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Conflicts of interest

J.K.B., W.D. and U.W. are employees of SwissRe: this company may benefit financially from increased insurability of HIV-positive lives.

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Ethical approval: Institutional review boards from each cohort approved analysis of routinely collected data to be used for research purposes. The ART Cohort Collaboration does not require ethical approval as all data is anonymised before transfer to ART-CC.

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