INTRODUCTION

It is well known that smoking is an important cardiovascular and cerebrovascular risk factor. Now, some studies have proved that smoking is not only a major harm to cardiovascular and cerebrovascular health, but also a major cause of chronic obstructive pulmonary disease. Smoking is associated with accelerated decline in lung function, increased mortality, and worsening symptoms in asthma and COPD (Boulet et al., 2006). There are also meta-analyses that show smokers have an increased risk of fracture (Kanis et al., 2005; Shen et al., 2015).

Vitamin D is one of the most important fat-soluble vitamins in the body, the most important of which are D2 and D3. As the main circulating form of vitamin D, 25(OH)D plays an important role in regulating cell proliferation and immune response, in addition to regulating serum calcium and phosphorus levels (Umar et al., 2018; Vanherweghen et al., 2017). Studies have found that inhibiting the production of cAMP protein induced by 1,25(OH) 2D3 can enhance
the growth of mycobacterium (Liu et al., 2007). A meta-analysis found that vitamin D deficiency increases the risk of ischemic stroke (Zhou et al., 2018). Vitamin D deficiency has also been related to type 2 diabetes (Berridge, 2017; Lucato et al., 2017) and may increase the risk of falls (Girgis et al., 2013). Vitamin D deficiency is now a major health problem in the real world.

Many studies have shown the effect of smoking on circulating 25-hydroxyvitamin D, although someone has considered that no statistically significant effect on circulating vitamin D levels (Stürmer et al., 2015). However, there is growing evidence of the negative effects of smoking on 25-hydroxyvitamin D and calcium metabolism (Cuomo et al., 2019; Cutillas-Marcos et al., 2012; Hermann et al., 2000; Kim et al., 2018; Richard et al., 2017). A study by Jääskeläinen et al. (2013) of 5,714 subjects (47% males) aged 30–79 years found that smokers had lower serum 25(OH)D concentrations than nonsmokers. Hermann et al. (2000) found that current smokers had significantly lower serum 25-OHD levels than nonsmokers. Although the decrease was only 6.8%, serum levels were negatively correlated with the number of cigarettes currently smoked per day ($r = −16; p = .003$). In addition, Thuesen et al. (2012) in a recent large population study showed that the odds ratios of severe vitamin D deficiency ($25$(OH)D < 10 ng/ml)/vitamin D deficiency ($25$(OH)D < 20 ng/ml) associated with daily smoking were 1.47 and 1.36, respectively. This cannot be explained by other confounding lifestyle factors. After adjustment for other confounders, the adverse effects of smoking on circulating vitamin D levels were still suggested (Brot et al., 1999).

There is no consensus on whether smoking causes damage to circulating vitamin D levels, which may have to do with differences in study design or participant characteristics (such as age, sex, weight, health status), and techniques (testing methods) between studies. In recent years, research has continued on whether low levels of vitamin D are common in smokers. To conduct the impact of smoking behavior on circulating vitamin D levels, we conducted a meta-analysis of published studies to draw conclusions and provide public advice on clinical and public health issues.

## 2 MATERIALS AND METHODS

### 2.1 Search strategy

Studies on smoking and circulating vitamin D published before 31 December 2020 were searched from the following databases: MEDLINE(PubMed), EMBASE, and Cochrane. The retrieval strategy was ((25-hydroxyvitamin D[ALL] OR 25(OH)D[ALL] OR bone marker[ALL] OR Parathyroid Hormone[ALL] OR PTH[ALL]) AND (smoking[ALL] OR cigarette[ALL] OR smoker[ALL])).

### 2.2 Selection criteria

All studies included in this meta-analysis met the following five criteria: (1) Diseases affecting vitamin D absorption or metabolism were excluded, (2) clear smoking status and circulating vitamin D data, (3) all subjects are 18 years or older, (4) reported the types of samples, and (5) compare the circulating vitamin D status of smokers with that of nonsmokers. Any of the following studies were excluded: participating in osteoporosis studies or not excluding diseases that affect vitamin D levels; data cannot be obtained, extracted, or documented (or if the author does not respond when attempting to contact); studies on minors; patients who use marijuana or drugs; and studies on vitamin D deficiency are ill-defined.

### 2.3 Data extraction

The following information and clinical characteristics of the participants were extracted from the included study: first author, year of publication, research type, detection method, sample type, study location, average age, body mass index (BMI), number of participants, number of smokers and nonsmokers, circulating vitamin D levels in smokers and nonsmokers (dichotomous variables studies were used to investigate vitamin D deficiency in smokers and nonsmokers), and use of vitamin D supplements.

### 2.4 Quality assessment

The quality assessment of the included studies used the Newcastle-Ottawa scale (NOS) based on cohort studies and case-control studies. We evaluated the article on the basis of selection, comparability, and outcome/exposure, with a possible total quality evaluation of 10. The Australian JBI Centre for Evidence-Based Health Care (2016) Quality Assessment Tool was used as the evaluation criteria for the cross-sectional study. All articles are evaluated by a reviewer.

### 2.5 Outcome measures

Our primary outcome measure was adult circulating vitamin D levels, compared between smokers and nonsmokers, and performed a subgroup analysis based on meta-regression results.

### 2.6 Statistical analyses

Continuous results were calculated using Cohen’s method as mean difference (SMD) and 95% confidence interval (CI), and the combined effect values of dichotomous variables were RR and its 95% confidence interval, using an inverse-variance model. $I^2$ statistics were used to test heterogeneity among studies. A value of less than 50% of $I^2$ is considered low heterogeneity. In statistics, when there is significant heterogeneity, the size of the mixed effect is calculated by using the random effect model. Otherwise, a fixed effect model is used. Suppose $p$ value < .05 or 95% CI without 0 (the 95% CI for the RR value does not include 1) was considered statistically significant. Based on the results of meta-regression, the main sources of
TABLE 1 Characteristics of included studies

![Table image]

| Characteristics of dichotomous variable studies |
|------------------------------------------------|
| Author                      | Year   | Type of experiment | Number of participants | Age (Mean) | BMI (Mean) | Detection method | Sample  |
|------------------------------|--------|--------------------|------------------------|------------|------------|------------------|---------|
| Sara Bianchi                | 2012   | Cross-sectional    | 185                    | 60         | 25         | CLIA             | Serum   |
| Aline Richard               | 2017   | Cohort study       | 204                    | 31.1       | 22.74      | Elecsys          | Serum   |
| Rune Tønnesen               | 2016   | Cross-sectional    | 700                    | 21.8       | 23.22      | CLIA             | Serum   |
| F. Okan                     | 2018   | Cross-sectional    | 72                     | 73.5       | 27.8       | Elecsys          | Serum   |
| Kuibao Li                   | 2016   | Cross-sectional    | 348                    | 62.4       | 26         | ELISA            | Serum   |
| Eneida Boteon Schmitt       | 2017   | Cross-sectional    | 463                    | 57.8       | 29.04      | CLIA             | Serum   |
| Louise Lind Schierbeck      | 2012   | Cohort study       | 2,013                  | 50         | 25.11      | RIA              | Serum   |
| Jaydip Ray Chaudhuri        | 2013   | Cross-sectional    | 150                    | 49.4       |            | CMIA             | Serum   |
| Sun Hea Kim                 | 2018   | Cross-sectional    | 2,687                  | 72.26      | 23.83      | Unclear          | Serum   |
| YaWen Lu                    | 2020   | Cross-sectional    | 1,798                  | 62         | 24.54      | CLIA             | Serum   |
| HaoWei Xu                   | 2020   | Cross-sectional    | 232                    | 65.5       | 23.34      | Unclear          | Serum   |
| Alessandro Cuomo            | 2019   | Cross-sectional    | 290                    | 47.8       | 26.2       | Unclear          | Serum   |
| Giovanni Targher            | 2006   | Cross-sectional    | 390                    | 57.7       | 28.5       | CLIA             | Serum   |

| Characteristics of continuous variable studies |
|------------------------------------------------|
| Author                          | Year   | Type of experiment | Number of participants | Age (Mean) | BMI (Mean) | Detection method | Sample |
|----------------------------------|--------|--------------------|------------------------|------------|------------|------------------|--------|
| Michael Stürmer                 | 2015   | Cross-sectional    | 146                    | 57         | 27         | Elecsys          | Plasma |
| A. P. Hermann                   | 1999   | Cohort study       | 2015                   | 50.6       |            | RIA              | Serum   |
| N. Marta Diaz-Gómez             | 2007   | Cohort study       | 61                     | 30.13      | 22.86      | RIA              | Serum   |
| Eugenia Cutillas-Marco          | 2012   | Cross-sectional    | 177                    | 47         | 23.84      | CLIA             | Serum   |
| Eva N. Kassi                    | 2014   | Cross-sectional    | 181                    | 34.69      | 25.94      | LC-MS/MS         | Serum   |
| A. Supervía                     | 2005   | Cross-sectional    | 74                     | 32.26      | 23.59      | RIA              | Serum   |
| Jennifer K. Mulligan            | 2014   | Case-control study | 21                     | 46.29      |            | ELISA            | Plasma  |
| Muhammad Shah Alam              | 2018   | Cross-sectional    | 50                     | 40.64      | 26.49      | CLIA             | Serum   |
| Klingberg, E                    | 2015   | Cross-sectional    | 540                    | 40.5       | 24.8       | CLIA             | Serum   |
| A Inkeri Lokki                  | 2018   | Case-control study | 359                    | 27.06      | 24.12      | CMIA             | Serum   |
| Rolf Jorde                      | 2019   | Cohort study       | 406                    | 51.9       | 27.8       | Unclear          | Serum   |

3 | RESULTS

3.1 | Characteristics of the chosen articles

Two thousand five hundred forty-one articles were retrieved from the database (Figure 1). After reviewing the summary, we selected 218 complete publications for further evaluation based on our inclusion criteria. Of these studies, 150 were excluded because patients with diseases that affect vitamin D absorption or metabolism were not excluded, and eight were excluded because the samples were not adult venous blood. There were 36 studies did not have the required data or did not contain sufficiently detailed data and that could not be obtained by contacting the authors. Finally, this meta-analysis included 24 articles (Alam et al., 2018; Bianchi et al., 2012; Chaudhuri et al., 2013; Cuomo et al., 2019; Cutillas-Marco et al., 2012; Díaz-Gómez et al., 2007; Hermann et al., 2000; Jorde et al., 2019; Kassi et al., 2015; Kim et al., 2018; Klingberg et al., 2015; Li et al., 2016; Lokki et al., 2020; Lu et al., 2020; Mulligan et al., 2014; Okan et al., 2020; Richard et al., 2017; Schierbeck et al., 2012; Schmitt et al., 2018; Stürmer et al., 2015; Supervia et al., 2006; Tonnesen et al., 2016; Xu et al., 2020) (Table 1). The 24 studies involved 11,340 participants: 1,399 smokers and 9,941 nonsmokers included in this meta-analysis (when quitters are present, they are classified as nonsmokers...
and compared with current smokers and current nonsmokers). The following vitamin D detection methods were used in 24 selected articles: electrochemiluminescence immunoassay, chemiluminescent immunoassay, LC-MS/MS, ELISA, RIA, and CMIA. Among them, two samples were plasma and the remaining 22 samples were serum. The specific characteristics of each study were shown in Table 1.

### 3.2 Quality assessment

The quality of the studies was evaluated according to NOS and JBI standards. The literature in this meta-analysis included five cohort studies, two case-control studies, and 17 cross-sectional studies. We judged study quality in terms of selection, comparability, and results/exposure of studies according to the Newcastle–Ottawa scale of observational studies (Table 2). The total quality scores of the seven articles (cohort and case-control studies) ranged from 4 to 9, with a possible total of 10 points. The authenticity evaluation results of the cross-sectional study according to the JBI quality evaluation tool are shown in Table 2. Due to strict inclusion and exclusion criteria to exclude the influence of certain diseases on circulating vitamin D levels, these studies had low selection bias. Since the purpose of this study was to explore the effects of smoking on circulating vitamin D levels in adults, the exposure factors were identified as smoking, and subject information was obtained using questionnaires or written self-reports; no study using smoking as an exposure factor was subject blind.
3.3 | Vitamin D deficiency or insufficiency in smokers and nonsmokers in the dichotomous variable studies

People who are current smokers are more likely to have circulating vitamin D deficiency or deficiency than nonsmokers. We obtained a 95% confidence interval combined effect amount (RR 1.11, 95% CI: 1.03–1.19, p < .001), p < .05 and 95% CI did not contain 1. The model is fixed, inverse-variance, I² = 75% (Figure 2). No source of heterogeneity was found in meta-regression. Sensitivity analysis suggested that study by Yawen Lu et al. was the main source of heterogeneity. When that study was removed, I² dropped to 38%. To better rule out the effect of vitamin D supplement use on the results, we analyzed the use of vitamin D supplements in subgroups. The use of vitamin D supplements was defined as 1, and nonuse of vitamin D supplements was defined as 0. The results are as follows (Figure 3): The combined effect size was available in people who did not receive vitamin D supplementation (RR 1.29, 95% CI: 1.09–1.53, p < .05), intragroup heterogeneity I² = 0 in this subgroup. A combined effect size was obtained when vitamin D supplements were known to be used (RR1.17, 95% CI: 1.06–1.28, p < .05), I² = 72.5% in this subgroup.

3.4 | Circulating vitamin D levels in smokers and nonsmokers in continuous variable studies

People who are current smokers had lower circulating vitamin D levels than nonsmokers. Due to the different vitamin D units provided in the literature, the combined effect amount was expressed by SMD. We obtained a 95% CI combined effect amount (SMD −0.24, 95% CI: −0.31 to −0.17, p < .001), p < .05 and 95% CI did not contain 0, which was statistically significant. I² = 70.1%. Due to the large heterogeneity, REML method was used to conduct univariate meta-regression analysis one by one with the number of subjects, type of experiment, year of publication, average age, test method, sample type, and use of vitamin D supplement as covariables, and the combined effect amount as dependent variables. The results are as follows (Table 3): Meta-regression results suggested that age was the main source of heterogeneity (Coef. 0.022, 95% CI: 1.009–1.037, p < .05). Consider that most women reach menopause around the age of 50. Bone metabolism in premenopause and postmenopause is affected by fluctuation of estrogen level, and circulating vitamin D level may change accordingly. We divided the study into three subgroups with mean age over 50, 40–50 years old (excluding 50), and under 40 years old (excluding 40) as the dividing line, and the results are as follows (Figure 4). In the over 50 group (group 1), we get a combined effect size (SMD-0.16, 95% CI: −0.29 to −0.03, p < .05), intragroup heterogeneity I² = 24.4%. In 40- to 50-year-old group (group 2), we get a combined effect size (SMD-0.49, 95% CI: −0.86 to −0.12, p < .05), intragroup heterogeneity I² = 21.4%. Under 40-year-old group (group 0) to get a combined effect size (SMD-0.57, 95% CI: −0.82 to −0.31, p < .05), intragroup heterogeneity I² = 56.8%.

3.5 | Publication bias

Begg’s test was used to examine the publication bias of the literature. Adjustment statistics of dichotomous variable studies Z = 0.18, p = .855, p > .05. Adjustment statistics of continuous variable studies Z = 1.4, p = .161, p > .05. There was no publication bias in the studies of dichotomous variables and continuous variables.

4 | DISCUSSION

This meta-analysis provides evidence of the negative effects of smoking on circulating vitamin D levels. These results summarized...
the findings from some small samples, the meta-analysis showed that smoking was associated with lower levels of circulating vitamin D, and the influence between smokers and nonsmokers and whether to use vitamin D supplements no correlation, because no matter whether to use vitamin D supplements, the results of the meta-analysis indicate current smokers were more likely than nonsmokers to circulating vitamin D deficiency or inadequate. In the group without vitamin D supplement, $I^2 = 0$, intragroup heterogeneity did not exist; while in the group with vitamin D supplement, $I^2 = 72.5\%$, intragroup heterogeneity was high. It may be related to the differences in the dose and dosage form of vitamin D taken by study participants, or it may be related to the small number of studies included in the group. In the study of continuous variables, we preliminarily obtained a result with high heterogeneity ($I^2 = 70.1\%$). After meta-regression, it was known that age was the main factor causing heterogeneity. We divided different age groups into different subgroups for analysis. We obtained two results with low heterogeneity in the over-50 group and the 40–50 group and one result with moderate heterogeneity in the under-40 group. Since most women of childbearing age are under the age of 40, we hypothesized that the source of heterogeneity might be the use of contraceptives by women of childbearing age, resulting in an impact on the results. Circulating vitamin D levels in smokers were lower than in nonsmokers at different ages, and this was particularly evident in the under 40 age group and on the surface that the effect of smoking on circulating vitamin D levels decreases with age. But studies have shown that smoking has the greatest effect on bones in older people, while no significant effect has been observed in any part of the bones in people under 40 (Ward & Klesges, 2001). The effect of smoking on bone metabolism is greater in the elderly than in the young, and the dose effect of smoking on bone has been shown in some literature (McCulloch et al., 1991; Ortego-Centeno et al., 1997; Välimäki et al., 1994). The same phenomenon has been seen in studies of smoking and circulating vitamin D levels (Rapuri et al., 2000). This is contrary to our observation that smoking has a greater negative effect on circulating vitamin D levels in young people. This may have something to do with the fact that older people take vitamin D supplements at a higher rate than younger people, although there is no clear evidence to support this conclusion. However, after the exclusion of vitamin D supplements, the RR for smoking-induced vitamin D deficiency increased roughly with age.

Many studies have explored the impact of smoking on vitamin D endocrine system (VDES). Previous studies found that smoking can affect vitamin D metabolism in many ways, including vitamin
D intake, synthesis, hydroxylation, and catabolism. Tobacco smoke contains substances such as polycyclic aromatic hydrocarbons, aldehydes, and DDT and is a carcinogen, neurotoxin, and endocrine disruptor (Diamanti-Kandarakis et al., 2009; Smith & Hansch, 2000). There are several hypotheses about the specific mechanism by which smoking lowers vitamin D levels.

Skin synthesis is the main source of vitamin D in the human body, and vitamin D synthesis is affected by skin aging. Smoking (Ernster et al., 1995) and ultraviolet radiation from the sun (Pillai et al., 2005) are considered important factors that contribute to skin aging in humans. Photosynthesis of precholecalciferol in human skin, that is, solar ultraviolet B photons with energy between 290 and 315 nanometer penetrates the skin under sunlight, making 7-dehydrogenated cholesterol (provitamin D3)(7-DHC) photolysis into precholecalciferol (provitamin D) (Figure 5). The precholecalciferol is thermodynamically unstable and requires isomerization to form the cholecalciferol. Once formed, cholecalciferol passes from the skin into the bloodstream, where it binds to vitamin D-binding proteins (Holick et al., 1980; Kira et al., 2003; MacLaughlin et al., 1982). Increased skin aging significantly reduces the skin's ability to convert 7-DHC to precholecalciferol (Holick, 1995). López Hernández et al. (1995) found evidence of accelerated skin aging among smokers, finding that smoking, sun exposure, and age are important factors that increase skin wrinkling, and skin wrinkles are a significant sign of skin aging. Matrix metalloproteinases (MMPs) induce photoaging, and there is ample evidence to emphasize the effect of smoking on skin aging by activating MMPs (Holick, 1995; Lahmann et al., 2001). This is consistent with our findings that in people who do not take vitamin D supplements, the risk of vitamin D deficiency in smokers increases with age. Smoking can also affect the expression of cytokines and inflammatory mediators. In the study of Tsutakawa et al. (2009), it was found that nicotine (0.35 mg kg⁻¹ day⁻¹) significantly enhanced the protein expression of cyclooxygenase 2(COX-2) and inducible nitric oxide synthase (iNOS), increased the release of pro-inflammatory mediators, and caused damage or delayed healing of skin and blood vessels.

On the other hand, the decrease in vitamin D may be related to the inhibition of PTH caused by smoking. Several studies we included (Cutillas-Marco et al., 2012; Díaz-Gómez et al., 2007; Supervia et al., 2006) detected parathyroid hormone levels in smokers and found that PTH levels were lower in smokers than in nonsmokers. Parathyroid hormone levels are mainly regulated by calcium ions. In the study by Need et al. (2002) (Schwarz et al., 1994), smokers had higher levels of ionized calcium in their serum, and small changes in ionized calcium levels caused rapid changes in parathyroid hormone secretion and synthesis. Several other hypothesized regulators, such as chromium-granin peptides and interleukin-8, may also be involved in regulating PTH secretion (Jorde et al., 2005). In

![Figure 2](image-url)  
**FIGURE 2** Vitamin D deficiency or insufficiency in adult smokers
addition, smoke may have a direct toxic effect on parathyroid cells. Jorde et al. (2005) hypothesized that there may be some substances in cigarette smoke that interact directly with calcium receptors, allowing smoking to enhance the degradation of parathyroid hormone in blood samples, leading to a decrease in parathyroid hormone levels without affecting calcium homeostasis. Studies have shown that low serum concentrations of VD and the dose–response pattern of smoking, that is, smoking for longer periods of time and smoking more, lead to lower levels of vitamin D (Hermann et al., 2000; Jiang et al., 2016). Although studies have suggested that current smokers have lower levels of parathyroid hormone than nonsmokers (Brot et al., 1999; Jorde et al., 2005, 2019; Landin-Wilhelmsen et al., 1995; Paik et al., 2010), it is not clear whether this dose relationship exists. But both parathyroid hormone and vitamin D levels are similar to those of never smokers after smoking cessation (Jorde et al., 2005). We conjectured that decreased vitamin D levels in smokers may be associated with decreased parathyroid hormone secretion, since PTH enhances the effect of renal 1-hydroxylase on vitamin D3 activation (Lips, 2001).

In addition, vitamin D3 synthesized in the skin and vitamin D2 extracted from food circulating in the body are first hydroxylated as they pass through the liver and then hydroxylated by the 1α-hydroxylase CYP27B1 as they pass through the kidneys to form active vitamin D1,25(OH)2D (Omdahl et al., 2002). Long-term smoking can lead to accumulation of heavy metals in the body. Tobacco plants contain lead and cadmium (Lugon-Moulin et al., 2006). The decrease in serum 1,25(OH)2D levels in smokers is related to the accumulation of cadmium in the kidneys (Brot et al., 1999; Kido et al., 1989). The mechanism of this process may be renal damage caused by cadmium and lead poisoning. In fact, the increase in lead and cadmium in the body can damage renal tubular function and glomerular function (Cooper, 2006).

Smoking can also lead to disorders in the catabolism of vitamin D. A polycyclic aromatic hydrocarbon, benzopyrene (BaP) in tobacco...
smoke increases the recruitment of CYP24A1 promoters by 1,25(OH)2D3-dependent VDR and retinoid X receptors, which promotes the decomposition of 1,25(OH)2D3 in human monocyte/macrophage-derived THP-1 cells and reduces vitamin D levels by enhancing induced expression of cytochrome P450 24A1 (Matsunawa et al., 2009). Another arene receptor (AhR) ligand, 2,3,7,8-tetrachlorodibenzo-p-dioxins, enhances CYP24A1 expression of 1,25(OH)2D3 in THP-1 cells. After treatment with AhR antagonists and protein synthesis inhibitors, the enhancement of CYP24A1 induced by BaP was inhibited, indicating that the role of BaP was mediated by AhR activation and de novo synthesis of proteins. Therefore, the activation of AhR by BaP can promote catabolism of vitamin D3 (Matsunawa et al., 2009).

Other researchers have found that SNP rs4809957 located in the three untranslated regions of CYP24A1 gene 20q13.2 interacts with smoking, and 1,25(OH)2D3 plays an antiproliferation role on vitamin D receptor-mediated human cancer cells (Dong et al., 2012). When smokers’ CYP24A1 genes were affected by smoke, lower vitamin D levels were linked to lung cancer risk.

Other studies have shown that tobacco alters the sense of smell and taste of food, leading to lower levels of vitamin D intake (Frye et al., 1990; Grunberg, 1982; Morabia et al., 2000; Redington, 1984). It has been reported that other peripheral tissues, including respiratory epithelial cells, also contain 1α-hydroxylase and may serve as a source of 1,25VD3 (Hansdottir et al., 2008). The researchers found that acrolein and extract (CSE) from tobacco smoke significantly reduced the conversion of 25(OH)D3 to 1,25(OH)2D3 in airway epithelial cells by downregulating the expression of CYP27B1 (CYP27B1 is the gene encoding the 25-hydroxyvitamin D3 1α-hydroxylase responsible for converting 25VD3 to 1,25(OH)2D3) (Mulligan et al., 2014). One of the articles we included was a study of patients with chronic rhinosinusitis (Mulligan et al., 2014), showed that smokers, regardless of whether they had chronic rhinosinusitis, had lower circulating vitamin D levels than nonsmokers.

This study is the first meta-analysis of circulating vitamin D levels in smokers. There have been many independent studies, but this is the first time we have combined statistics on this subject. Because diseases affecting vitamin D absorption or metabolism were excluded in advance, the results of this meta-analysis were not affected by adverse or beneficial effects of disease on vitamin D sources and metabolism. Confounding factors such as estrogen replacement therapy may partially mask the effect on women, and oral contraceptives or estrogen may prevent bone loss, increasing error

### FIGURE 4

Vitamin D levels of smokers at different ages
variability. So we also excluded menopausal women who were on hormone replacement therapy. However, for women of childbearing age using oral contraceptives, it is unclear whether these women of childbearing age using contraceptives have been excluded since no such information was provided in the article, which may be one of the sources of heterogeneity in the subgroup of this population. In addition, several limitations of our study must be noted. Since vitamin D samples were taken from serum and plasma, this may lead to interstudy heterogeneity. However, meta-regression showed that the sample type was not the main source of interstudy heterogeneity, which had little impact on the results, and the few studies with the sample type of plasma were not suitable for further subgroup analysis. More research is needed to confirm whether vitamin D levels in smokers are reduced in serum and plasma, respectively. Our study did not analyze the effect of exercise on VD levels. Some studies have suggested that physical activity has an effect on vitamin D levels, but studies until now have not established a direct relationship between physical activity and VD (Bell et al., 1988; Jacques et al., 1997; Klausen et al., 1993; Lucas et al., 2005; Ma Moun et al., 2005; Scragg et al., 1992, 1995), or that sun exposure is a confounding factor, and there are conflicting results regarding the effect of physical activity on 25(OH)D metabolism. It was found that there was no significant relationship between physical activity and 25(OH)D after controlling for sunshine duration. In addition, since the included studies came from different countries and regions, and the included studies were mainly retrospective studies, the time, season, and latitude of blood sample extraction of subjects could not be unified, and these confounding factors might affect the results and produce heterogeneity among studies. However, in a single study, the subjects had the same season, latitude, and blood sampling time, and there were relevant measures to exclude the influence of confounding factors, so the study results were less affected by these factors. The association between smoking and circulating vitamin D levels is unlikely to be accidental, as estimates come from a large mixed sample (more than 11,340 participants) and these studies are heterogeneous in terms of several participant and methodological characteristics. Due to the exclusion of some diseases that affect vitamin D absorption or metabolism, there is inevitably a selection bias, and since all studies rely on self-reporting, there may be bias in determining smoking exposure.

Smoking-related vitamin D deficiency (VDD) may pose an even greater public health problem in our aging society. Given the cumulative and dose-dependent effects of smoking on VDD and its association with the vitamin D-PTH axis, an increase in smoking among these young populations is likely to lead to a significant increase in the future public health burden of osteoporosis or other cardiovascular, pulmonary, infectious, and immune-related diseases. Smoking has a significant adverse effect on circulating vitamin D levels, but vitamin D levels can be restored to nonsmoker levels after quitting smoking. In view of the adverse effects of smoking on vitamin D-PTH axis health and its associated public effects on bone metabolism and other systemic health, the prevention of VDD and smoking cessation must be emphasized.

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Studies included in this paper are all published human studies. None of the authors participated in any clinical study and therefore did not sign informed consent.
CONFLICT OF INTEREST
The authors declare that there is no conflict of interest.

AUTHOR CONTRIBUTIONS
Lu Yang: Formal analysis (lead); Writing-original draft (equal). Hang Zhao: Formal analysis (supporting); Methodology (equal). Ke Liu: Investigation (equal). Yichao Wang: Investigation (equal). Tianitian Sun: Investigation (equal). Qianqian Liu: Investigation (equal). Shuchun Chen: Writing-review & editing (equal). Luping Ren: Writing-review & editing (equal).

DATA AVAILABILITY STATEMENT
The data that supports the findings of this study are available in this article. No additional supplementary material is provided.

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