Application of Response Surface Methodology (RSM) for Optimization of Anti-Obesity Effect in Fermented Milk by Lactobacillus plantarum Q180

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Abstract
Obesity, a condition in which an abnormally large amount of fat is stored in adipose tissue, causing an increase in body weight, has become a major public health concern worldwide. The purpose of this study was to optimize the process for fermented milk for the production of a functional product with an anti-obesity effect by using Lactobacillus plantarum Q180 isolated from human feces. We used a 3-factor, 3-level central composite design (CCD) combined with the response surface methodology (RSM). Concentration of skim milk powder (%), incubation temperature (°C), and incubation time (h) were used as the independent factors, whereas pH (pH), anti-lipase activity (%), and anti-adipogenetic activity (%) were used as the dependent factors. The optimal conditions of fermented milk for the highest anti-lipase and anti-adipogenetic activity with pH 4.4 were the 9.5% of skim milk powder, 37°C of incubation temperature, 28 h of incubation time. In the fermentation condition, the predicted values of pH, anti-lipase activity and anti-adipogenetic activity were 4.47, 55.55, and 20.48%, respectively. However, the actual values of pH, anti-lipase activity and anti-adipogenetic activity were 4.50, 52.86, and 19.25%, respectively. These results demonstrate that 9.5% of skim milk powder and incubation at 37°C for 28 h were the optimum conditions for producing functional fermented milk with an anti-obesity effect.

Keywords: Lactobacillus plantarum, optimization, anti-lipase activity, anti-adipogenetic activity

Introduction
Obesity is becoming increasingly prevalent among adults, adolescents and children and has become a major public health concern worldwide (Yanovski and Yanovski, 2002). Indeed, obesity is closely related with several metabolic syndromes such as hypertension, diabetes, hyperlipidemia, and arteriosclerosis (Tanida et al., 2008). For these reasons, numerous people have an interest in this issue.

Lactic acid bacteria (LAB) possess special physiological activities and are generally regarded as safe (GRAS). LAB have been widely used in a number of fermented foods, particularly in the production of dairy and vegetable products with functional and probiotic properties (Karaean et al., 2010; Leroy and Vuyst, 2004). As regards their use as probiotics, LAB are reported to have various beneficial effects on the health of hosts once consumed in adequate amounts. These effects include the modulation of immune responses (Salminen et al., 2002), and anticarcinogenic and anti-oxidative activities (Choi et al., 2006). In addition to these effects, certain LAB have been found to be effective in regulating adipose tissue in overweight adults (Kadooka et al., 2010) as well as in a diet-induced obese animal model (Takeamura et al., 2010).

Individual LAB has a specific fermentation profile, such as the ability to form functional substances and to produce acid. Thus, taking the profiles of LAB into consideration is a significant factor when it is used as a starter in the production of fermented foods (Komatsuzaki et al., 2005). The response surface methodology (RSM), which was first described by Box and Wilson (Box and Wilson, 1951), is a collection of statistical and mathematical techniques. It is based on the fit of a polynomial equation to experimental data (Bezerra et al., 2008). Because RSM is an efficient experimental strategy for seeking optimal conditions for a multivariable system, it has been successfully employed in optimizing the culture conditions (Box et al., 1978).

The aim of this study is to optimize the fermentative parameters in order to apply them to functional food products which have an anti-obesity effect.

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Materials and Methods

Bacterial strains

A LAB strain having an anti-obesity effect, namely, *L. plantarum* Q180, was isolated from feces of healthy adults. In our previous study, *L. plantarum* Q180 was found to have lipase inhibitory activity of 83.61±2.32% and to inhibit the adipocyte differentiation of 3T3-L1 cells (14.63 ±1.37%) at a concentration of 100 µg/mL (Park et al., 2014). The strain was incubated in Lactobacilli MRS broth (Difco, USA) as the growth medium at 37°C for 18 h.

Anti-adipogenetic activity

Cell line and cell culture

3T3-L1 cells were cultured as described by Hemati et al. (1997). The 3T3-L1 cells were obtained from the American Type Culture Collection (ATCC, USA) and cultured in Dulbecco’s modified Eagle’s medium (DMEM; Gibco, USA) containing a high glucose content supplemented with 10% bovine calf serum (BCS; Gibco, USA) and 1% penicillin/streptomycin (Sigma, USA) at 37°C in a humidified 5% CO2 atmosphere. To induce differentiation, 2-d post-confluent cells (0 d) were stimulated for 2 d with an adipocyte differentiation cocktail medium containing 5 mM 3-isobutyl-1-methylxanthine (IBMX; Sigma, USA), 1 mM dexamethasone (Dex; Sigma, USA), and 5 g/mL of insulin (Sigma, USA) in DMEM supplemented with 10% fetal bovine serum (FBS; Gibco, USA) and 1% penicillin/streptomycin. On 2 d, the medium was replaced with DMEM containing 10% FBS, 1% penicillin/streptomycin, and 5 g/mL of insulin, and incubated for 2 d, followed by culturing with DMEM containing 10% FBS and 1% penicillin/streptomycin for an additional 4 d (8 d), at the end of which more than 90% of the cells were mature adipocytes with accumulated fat droplets.

Cell viability

Cell viability was assessed by the MTT (3-(4,5-Dimethyl-2-thiazolyl)-2,5-diphenyl-2H-tetrazolium bromide) assay. The MTT assay was performed according to the modified method of Mosmann (1983). The 3T3-L1 pre-adipocytes were placed in 96-well microliter plates at a density of 16×10² cells/well. After 24 h incubation, the culture medium was replaced by 100 µL of serial dilutions (10, 100, 1000 mg/mL) of the sample, and the cells were incubated for 24 h. After incubation, 20 µL of sterile filtered MTT solution (5 mg/mL) in PBS (PBS, 0.85% NaCl, 2.68 mM KCl, 10 mM Na2HPO4 and 1.76 mM KH2PO4 were dissolved in distilled water, pH 7.4) was added to each well. Unreacted dye was removed after 4 h incubation. Insoluble formazan crystals were dissolved in 100 µL/well of dimethyl sulfoxide (DMSO) and measured spectrophotometrically in an ELISA reader (BioTek, USA) at 550 nm (sample A). The non-treated cell was also dissolved in 100 µL/well of DMSO and the absorbance was recorded at 550 nm (control A). The percent viability was expressed using the following formula:

\[
\text{Cell viability} (\%) = 100 \times \left(\frac{\text{Absorbance of sample A} - \text{Absorbance of control A}}{\text{Absorbance of control A}}\right) \times 100
\]

Sample preparation and treatment *L. plantarum* Q180

*L. plantarum* Q180 was incubated at 37°C for 18 h in MRS broth. All of the purified strains were kept at 70°C until use. After culturing the L. plantarum Q180, all of the strains were harvested in a refrigerated centrifuge (1,500 g for 15 min at 4°C) and washed three times with distilled water to remove any remaining MRS broth. The washed *L. plantarum* Q180 was freeze-dried and re-suspended in distilled water at a concentration of 10 mg/mL and homogenized for 50 sec followed by 1 min of rest (repeated 3 times) using a sonicator. The 3T3-L1 cells were treated with 100 g/mL of the sample. The concentration of the sample was determined according to the result of the MTT assay.

Oil red O staining of 3T3-L1 adipocyte

Intracellular lipid accumulation was measured using oil red O (Sigma, USA). Oil red O staining of 3T3-L1 cells was performed using a modified version of the method described by Ramirez-Zacarias et al. (1992). 3T3-L1 cells were washed with PBS twice, fixed with 10% formaldehyde/PBS at 4°C for 1 h, and stained with filtered oil red O solution (stock solution: 3.5 mg/mL in isopropanol; working solution: 60% oil red O stock solution and 40% distilled water) at room temperature for 30 min. The quantification of lipid accumulation was achieved by the oil red O obtained from stained cells with isopropyl alcohol and measured spectrophotometrically at 520 nm. The material stained with oil red O was expressed on a per cell basis using the number of cells determined from similar plates. The percentage of the material stained with oil red O relative to the control wells containing the cell culture medium without compounds was calculated as 520 nm (Q180)/520 nm (control) × 100.
Anti-lipase activity
The method of determining lipase activity proposed by Lee et al. (1993) was modified. Pancreatic lipase activity was measured using porcine pancreatic lipase (Sigma, USA). 0.1 mg/mL of a sample solution dissolved in water, 0.167 mM p-Nitrophenylpalmitate (PNP; Sigma, USA) solution and 0.061 M (pH 8.5) Tris-HCl buffer were mixed in the well of a plate, and 0.3 mg/mL of the lipase solution was then added to start the enzyme reaction. After incubation at 25°C for 10 min, its absorbance was measured at 405 nm.

Experimental design
To optimize the fermentative condition of L. plantarum Q180, concentration of skim milk powder (%), incubation temperature (°C), and incubation time (h) were used as the independent factors. In this design there are three experimental levels: -1, 0, 1. The range and center point values of the three independent factors were chosen after a series of preliminary single factor experiments (Table 1). pH (pH, Y₁), anti-lipase activity (% Y₂), and anti-adipogenetic activity (pH, Y₃) were selected as the dependent factors.

Response surface methodology
The central composite design (CCD) described by Box and Wilson (1951) was adopted for the optimization of the anti-obesity activity of L. plantarum Q180. The CCD in the experimental design consisted of 2³ factorial points, six axial points (α = 2), and three replicates of the central point (Table 2). Experimental runs were randomized in order to minimize the effects of unexpected variabilities in the observed responses.

Analysis of the data
The statistical analysis of the data and the multiple response optimizations were calculated by the desirability function of MINITAB statistical software (Version 13, Minitab Inc., USA). The statistical analysis was performed to fit the following quadratic polynomial equation:

\[
Y = \beta_0 + \sum_{i=1}^{3} \beta_iX_i + \sum_{i=1}^{3} \beta_{ii}X_i^2 + \sum_{i=1, j=1}^{3} \beta_{ij}X_iX_j
\]

Where Y represents the dependent variables (pH, anti-lipase activity and anti-adipogenetic activity), \( \beta \) is constant, \( \beta_i, \beta_{ii}, \beta_{ij} \) are regression coefficients, and \( X_i, X_j \) are levels of the independent variables. Multiple response optimizations were performed to search for the condition

Table 1. Independent variables and their levels in the 3-factor, 3-level central composite rotatable design optimizing the incubation condition of L. plantarum Q180

| Independent variables       | Symbol | Level |
|-----------------------------|--------|-------|
| Skim milk powder (%)        | X₁     | -1 0 1|
| Incubation temp. (°C)       | X₂     | 9 10 11|
| Incubation time (h)         | X₃     | 34 37 40|

Table 2. Central composite design and responses of dependent variables for fermented milk with Lactobacillus plantarum Q180 to independent variables

| Run No. | X₁  | X₂  | X₃  | Y₁  | Y₂  | Y₃  |
|---------|-----|-----|-----|-----|-----|-----|
| 1       | -1  | -1  | -1  | 5.56| 22.92| 19.17|
| 2       | 1   | -1  | -1  | 5.87| 11.94| 2.39 |
| 3       | -1  | 1   | -1  | 4.70| 46.49| 13.73|
| 4       | 1   | 1   | -1  | 4.72| 5.41 | 5.94 |
| 5       | -1  | -1  | 1   | 5.38| 55.40| 10.29|
| 6       | 1   | -1  | 1   | 5.44| 51.05| -4.02|
| 7       | -1  | 1   | 1   | 4.14| 67.44| 12.28|
| 8       | 1   | 1   | 1   | 4.29| 33.47| 5.58 |
| 9       | -1.68179| 0   | 0   | 4.43| 48.65| 26.78|
| 10      | 1.68179| 0   | 0   | 4.63| 56.54| -2.57|
| 11      | 0   | -1.68179| 0   | 6.02| 5.58 | 13.91|
| 12      | 0   | 1.68179| 0   | 4.72| 47.20| 5.76 |
| 13      | 0   | 0   | -1.68179| 5.40| 21.75| 30.04|
| 14      | 0   | 0   | 1.68179| 4.23| 60.71| 10.11|
| 15      | 0   | 0   | 0   | 4.55| 56.60| 18.44|
| 16      | 0   | 0   | 0   | 4.43| 53.75| 16.45|
| 17      | 0   | 0   | 0   | 4.40| 56.54| 15.00|

X₁: skim milk powder (%), X₂: incubation temp. (°C), X₃: incubation time (h); Y₁: pH, Y₂: anti-lipase activity (%), Y₃: anti-adipogenetic activity (%).
that could simultaneously satisfy the three dependent variables \(Y_1, Y_2\), and \(Y_3\). The response surface plots were developed using Maple software (Maple 7, Waterloo Maple Inc., Canada), and represented a function of two independent variables, while keeping the another independent variables at the optimal values.

Results and Discussion

Diagnostic checking of the fitted models

The pH, anti-lipase activity, and anti-adipogenetic activity were measured by the seventeen fermentation conditions (Table 2 and Fig. 1). MINITAB statistical software was employed to fit the quadratic polynomial equation to the experimental data. All the coefficients of linear \((X_1, X_2, X_3)\), square \((X_{11}, X_{22}, X_{33})\) and interaction \((X_{12}, X_{13}, X_{23})\) were calculated with a \(t\)-statistic to determine their significance. The estimated coefficients of each model are presented in Table 3. As a result of the significance test, in the case of \(Y_1\) (pH), \(X_3\) and \(X_3\) were found to be lower than the significance level \((p\)-value\) of 0.05 in the first-order term, and are thus statistically significant, and exercised great influence on the dependent variables. The cross-terms were not statistically significant, except for \(X_2X_3\), and \(X_3X_3\). In the case of \(Y_2\) (anti-lipase activity, \%), \(X_3\) was significant and exercised a great influence on the dependent variables. It was shown that the cross-terms were not statistically significant, except for \(X_2X_2\). In the case of \(Y_3\) (anti-adipogenetic activity, \%), \(X_3\) showed significance and had a great influence on the dependent variables, and none of the cross-terms was statistically significant. The reaction equation obtained on the basis of the above results is shown in Table 4. A proper second-order polynomial expression model was obtained on the basis of the results of the response surface analysis. The coefficients of determination \((R^2)\) for \(Y_1, Y_2\), and
were 0.963, 0.935 and 0.810, respectively, which indicates that the model is suitable to represent the real relationships among the selected reaction parameters. The values of $R^2$ for all models were extremely high for the response surface and significant at $p=0.00$. The reason why the values obtained for $R^2$ are quite high is that the model is well-fitted to the data.

| Variable and interaction | $Y_1$ Coefficient | $p$ value | $Y_2$ Coefficient | $p$ value | $Y_3$ Coefficient | $p$ value |
|--------------------------|-------------------|-----------|-------------------|-----------|-------------------|-----------|
| Intercept                | 4.45481           | 0.000     | 7.807             | 0.000     | 4.684             | 0.002     |
| $X_1$                    | 0.06417           | 0.216     | $-1.682$          | 0.136     | $-4.276$          | 0.004     |
| $X_2$                    | $-0.48227$        | 0.000     | 1.778             | 0.119     | 0.304             | 0.974     |
| $X_3$                    | $-0.26124$        | 0.001     | 4.061             | 0.005     | $-1.965$          | 0.090     |
| $X_1 X_2$                | 0.04262           | 0.439     | $-0.433$          | 0.678     | $-1.593$          | 0.155     |
| $X_1 X_3$                | 0.33961           | 0.000     | 2.941             | 0.022     | 0.034             | 0.978     |
| $X_2 X_3$                | 0.14338           | 0.028     | $-1.521$          | 0.172     | $-0.993$          | 0.270     |
| $-0.02500$              | 0.679             | 0.001     | $-1.703$          | 0.132     | 1.198             | 0.083     |
| $X_3$                    | $-0.01500$        | 0.815     | 0.392             | 0.707     | 0.467             | 0.655     |
| $-0.04750$              | 0.466             | 0.001     | $-0.644$          | 0.540     | 0.487             | 0.641     |

$X_1$: skim milk powder (%), $X_2$: incubation temp. (°C), $X_3$: incubation time (h); $Y_1$: pH, $Y_2$: anti-lipase activity (%), $Y_3$: anti-adipogenetic activity (%).

| Responses Quadratic polynomial model | $R^2$ | $p$-value |
|-------------------------------------|-------|-----------|
| $Y_1$ $Y_1=4.45481+0.06417X_1-0.48227X_2-0.26124X_3+0.04262X_1 X_2+0.33961X_1 X_3-0.14338X_1 X_3-0.25000X_2 X_3-0.02500X_1 X_2-0.01500X_1 X_3-0.04750X_2 X_3$ | 0.963 | 0.000     |
| $Y_2$ $Y_2=55.784-5.646X_1+5.967X_2+13.628X_3-1.599X_1-10.864X_2-5.617X_3-7.466X_1 X_3+1.717X_2 X_3-2.822X_3 X_3$ | 0.835 | 0.000     |
| $Y_3$ $Y_3=17.0408-7.3044X_1+0.0573X_2+3.3572X_3-2.9947X_1 X_2+3.7956X_2 X_3-0.1755X_1 X_3+2.6731X_1 X_3+1.0420X_2 X_3+1.0874X_3 X_3$ | 0.810 | 0.002     |

$X_1$: skim milk powder (%), $X_2$: incubation temp. (°C), $X_3$: incubation time (h); $Y_1$: pH, $Y_2$: anti-lipase activity (%), $Y_3$: anti-adipogenetic activity (%).

| Source of variation | DF | SS | MS | F-value | $p$-value |
|---------------------|----|----|----|---------|-----------|
| $Y_1$               |    |    |    |         |           |
| Main effects        | 9  | 5.57201 | 0.61911 | 20.39 | 0.000     |
| Linear              | 3  | 4.16466 | 1.38222 | 45.72 | 0.000     |
| Square              | 3  | 1.38249 | 0.46083 | 15.18 | 0.002     |
| 2-way interactions  | 3  | 0.02485 | 0.00828 | 0.27  | 0.843     |
| Residual error      | 7  | 0.21254 | 0.03036 |       |           |
| Lack of fit         | 5  | 0.19994 | 0.03999 | 6.35  | 0.042     |
| Total               | 16 | 5.78455 |          |       |           |
| $Y_2$               |    |    |    |         |           |
| Main effects        | 9  | 5.545.00 | 0.60600 | 3.94  | 0.042     |
| Linear              | 3  | 3.457.83 | 1.15262 | 7.50  | 0.014     |
| Square              | 3  | 1.462.97 | 0.48766 | 3.17  | 0.094     |
| 2-way interactions  | 3  | 0.533.17 | 1.77722 | 1.16  | 0.392     |
| Residual error      | 7  | 1.076.48 | 153.78  |       |           |
| Lack of fit         | 5  | 1.071.19 | 214.24  | 280.98| 0.012     |
| Total               | 16 | 6.530.48 |          |       |           |
| $Y_3$               |    |    |    |         |           |
| Main effects        | 9  | 1.118.26 | 132.029 | 3.31  | 0.064     |
| Linear              | 3  | 882.63  | 294.210 | 7.38  | 0.014     |
| Square              | 3  | 2.30.32  | 76.774  | 1.93  | 0.214     |
| 2-way interactions  | 3  | 0.753.17 | 25.103  | 0.63  | 0.619     |
| Residual error      | 7  | 2.79.01  | 39.859  |       |           |
| Lack of fit         | 5  | 2.73.03  | 54.607  | 18.27 | 0.023     |
| Total               | 16 | 14.672.77 |          |       |           |

DF, Degrees of freedom; SS, sum of squares; MS, Mean square (MS=SS/DF).

$Y_1$: pH, $Y_2$: anti-lipase activity (%), $Y_3$: anti-adipogenetic activity (%).

$Y_3$ were 0.963, 0.935 and 0.810, respectively, which indicates that the model is suitable to represent the real relationships among the selected reaction parameters. The values of $R^2$ for all models were extremely high for the response surface and significant at $p=0.00$. The reason why the values obtained for $R^2$ are quite high is that the model is well-fitted to the data.
Experimental design was based on an adequately performed preliminary test.

**Analysis of variance**

The statistical significance of the quadratic polynomial model equation was evaluated by conducting an analysis of variance (ANOVA). Table 5 shows the ANOVA for the models that explain the response of the three dependent variables, \( Y_1 \) (pH), \( Y_2 \) (anti-lipase activity) and \( Y_3 \) (anti-adipogenetic activity). The square terms and 2-way interaction terms for the dependent variables (\( Y_1 \), \( Y_2 \) and \( Y_3 \)) were not significant, except for the square terms of \( Y_1 \) (\( p=0.002 \) and \( p=0.843 \), \( Y_2 \); \( p=0.094 \) and \( p=0.392 \), \( Y_3 \); \( p=0.214 \) and \( p=0.619 \), respectively) at the 95% probability level (\( p<0.05 \)), whereas the linear term and total regression model were significant at the 95% probability level, except for the total regression model of \( Y_3 \). In this design, the data were highly influenced by the linear term. As regards the results of the lack-of-fit test, which indicates the fitness of the model, all the dependent variables were significant at the 95% probability level.

**Conditions for optimum responses**

In order to find the condition for optimum responses, the desirability function of the MINITAB statistical software was used. Optimal conditions included the coded and un-coded values of each dependent variable (\( Y_1 \), \( Y_2 \) and \( Y_3 \)), which are shown in Table 6. The results of performing the optimization of fermented milk each under conditions showing the highest anti-lipase activity and anti-adipogenetic activity and targeting the pH 4.4 showed that the critical values were different in all data. However, the result of the optimization satisfying both conditions at the same time showed that the coded values of the independent variables were the concentration of skim milk powder, \( X_1=-0.4572 \); incubation temperature, \( X_2=0.0548 \); and incubation time, \( X_3=-0.2134 \); respectively. The actual values of the independent variables against the coded values were \( X_1=9.5\% \), \( X_2=37^\circ C \) and \( X_3=28 \) h. The predicted values of multiple response optimal conditions were \( Y_1=\text{pH 4.47}, Y_2=55.55\% \) and \( Y_3=20.48\% \).

**Response surface plots and the effect of factors**

Fig. 2 shows the estimated response function and the effect of the independent variables (\( X_1 \), \( X_2 \) and \( X_3 \)) and dependent variables (\( Y_1 \), \( Y_2 \) and \( Y_3 \)). The response surface plot presents the interrelationship between two independent variables and one dependent variable, while keeping another independent variable at the optimal values. It is considered that two factors, i.e. incubation temperature and incubation time, affect anti-lipase activity (\( Y_2 \)) among the fermentation conditions for fermented milk with an excellent anti-obesity effect, and that all three independent variables affect the pH (\( Y_1 \)) and anti-adipogenetic activity (\( Y_3 \)).

**Verification of predicted values**

Verification experiments were conducted under optimal conditions (concentration of skim milk powder = 9.5\%, incubation temperature = 37\(^\circ C\), and incubation time = 28 h) to compare the predicted values and the actual values of the dependent variables (Table 7). The actual values, which were repeated three times, were \( \text{pH} = 4.50 \), anti-lipase activity = 52.86\%, and anti-adipogenetic activity = 19.25\% against the predicted values of \( \text{pH} = 4.47 \), anti-lipase activity = 55.55\%, and anti-adipogenetic activity = 20.48\%.

**Table 6. Optimal conditions of pH and GABA concentrations**

| Dependent | Independent variables | Critical value | Predicted value | Stationary point |
|-----------|----------------------|----------------|-----------------|-----------------|
| \( Y_1 \)  | \( X_1 \) \( X_2 \) \( X_3 \) | 0 0 0.2419 | 10 37 33.0399 | 4.4 Target |
| \( Y_2 \)  | \( X_1 \) \( X_2 \) \( X_3 \) | -1.6818 1.4017 0.6063 | 8.3182 41.2051 36.8661 | 67.42 Maximum |
| \( Y_3 \)  | \( X_1 \) \( X_2 \) \( X_3 \) | -1.4677 -0.5395 -1.6818 | 8.4677 65.3815 12.8412 | 31 Maximum |
| Multiple response optimization | \( X_1 \) \( X_2 \) \( X_3 \) | -0.4572 0.0548 -0.2134 | 9.5428 37.1644 28.2593 | |

\( X_1 \): skim milk powder (%), \( X_2 \): incubation temp. (\(^\circ C\)), \( X_3 \): incubation time (h); \( Y_1 \): pH, \( Y_2 \): anti-lipase activity (%), \( Y_3 \): anti-adipogenetic activity (%).
Both the actual values and the predicted values almost coincided with each other. According to Park et al. (2011), the lipid content in differentiated cells decreased by 11±3.6% when treated with 0.01% of *L. plantarum* KY1032, a strain isolated from kimchi. Therefore, the estimated response surface model has an excellent anti-obesity effect and can be adapted to optimize the production of functional fermented milk with an anti-obesity effect obtained from *L. plantarum* Q180.

### Conclusion

We investigated the optimum condition for producing functional fermented milk with an anti-obesity effect by using *L. plantarum* Q180. We used a 3-factor, 3-level CCD combined with RSM. Concentrations of skim milk powder (%), incubation temperature (°C), and incubation time (h) were used as the independent factors, while pH (pH, Y₁), anti-lipase activity (%, Y₂) and anti-adipogenetic activity (%, Y₃) were used as the dependent factors. The optimal conditions of fermented milk for the highest anti-lipase and anti-adipogenetic activity with pH 4.4 were the 9.5% of skim milk powder, 37°C of incubation temperature, 28 h of incubation time. In the fermentation condition, the predicted values of pH, anti-lipase activity and anti-adipogenetic activity were 4.47, 55.55, and 20.48%, respectively. However, the actual values of pH, anti-lipase activity and anti-adipogenetic activity were 4.50, 52.86, and 19.25%, respectively. These results demonstrate that 9.5% of skim milk powder and incubation at 37°C for 28 h were the optimum conditions for producing functional fermented milk with an anti-obesity effect.

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### Table 7. Predicted results of verification under optimized conditions

| Dependent | Predicted value | Experimental value |
|-----------|----------------|--------------------|
| Y₁(pH)    | 4.47           | 4.50±0.03          |
| Y₂(lipase) | 55.55          | 52.86±0.86         |
| Y₃(adipose) | 20.48          | 19.25±0.53         |

All values are mean±standard deviation of three replicates.
Agriculture, Food and Rural Affairs.

References

1. Bezerra, M. A., Santelli, R. E., Oliveira, E. P., Villar, L. S., and Escaleira, L. A. (2008) Response surfage methodology (RSM) as a tool for optimization in analytical chemistry. Talanta 76, 965-977.
2. Box, G. E. P., Hunter, W. G., and Hunter, J. S. (1978) Statistics for experimenters. John Wiley and Sons, New York, 341-418.
3. Box, G. E. P. and Wilson, K. B. (1951) On the experimental attainment of optimum condition. J. Roy. Statist. Soc. Ser. B. 13, 1-38.
4. Choi, S. S., Kim, Y., Han, K. S., You, S., Oh, S., and Kim, S. H. (2006) Effects of Lactobacillus strains on cancer cell proliferation and oxidative stress in vitro. Lett. Appl. Microbiol. 42, 452-458.
5. Hemati, N., Ross, S. E., Erickson, R. L, Groblewski, G. E., and MacDuygald, O. A. (1997) Signaling pathways through which insulin regulates CCAAT/enhancer binding protein-α (C/EBP-α) phosphorylation and gene expression in 3T3-L1 adipocytes. Forrelation with CLUT4 gene expression. J. Biol. Chem. 272, 25913-25919.
6. Kadooka, Y., Sato, M., Imaizumi, K., Ogawa, A., Ikuyama, K., Akai, Y., Okano, M., Kogoshima, M., and Tsuchida, T. (2010) Regulation of abdominal adiposity by probiotics (Lactobacillus gasseri SBT2055) in adults with obese tendencies in a randomized controlled trial. Eur. J. Clin. Nutr. 64, 636-643.
7. Karahan, A. G., Kilic, G. B., Kart, A., Aloglu, H. S., Oner, Z., Aydemin, S., Erkus, O., and Harsa, S. (2010) Genotypic identification or some lactic acid bacteria by amplified fragment length polymorphism analysis and investigation of their potential usage as starter culture combinations in Beyaz cheese manufacture. J. Dairy Chem. 93, 1-11.
8. Komatsu, N., Shima, J., Kawamoto, S., Momose, H., and Kimura, T. (2005) Production of γ-aminobutyric acid (GABA) by Lactobacillus paracasei isolated from traditional fermented foods. Food Microbiol. 22, 497-504.
9. Lee, Y. P., Chung G. H., and Rhee, J. S. (1993) Purification and characterization of Pseudomonas fluorescens SIK W1 lipase expressed in Escherichia coli. Biochim. Biophys. Acta. 1169, 156-164.
10. Leroy, F. and Vuyst, L. D. (2004) Lactic acid bacteria as functional starter cultures for the food fermentation industry. Trends Food Sci. Technol. 15, 17-78.
11. Mosmann T. (1983) Rapid colorimetric assay for cellular growth and survival: Application to proliferation and cytotoxicity assays. J. Immunol. Methods 62, 55-63.
12. Park, D. Y., Ahn, Y. T., Huh, C. S., Jeon, S. M., and Choi, M. S. (2011) The inhibitory effect of Lactobacillus plantarum KY1032 cell extract on the adipogenesis of 3T3-L1 cells. J. Med. Food. 10, 670-675.
13. Park, S. Y., Cho, S. A., Kim, S. H., and Lim, S. D. (2014) Physiological characteristics and anti-obesity effect of Lactobacillus plantarum Q180 isolated from feces. Korean J. Food Sci. An. 34, 645-653.
14. Ramirez-Zacarias, J. L., Castro-Munozledo, F., and Kuri-Harcuch, W. (1992) Quantitation of adipose conversion and triglycerides by staining intracytoplasmic lipids with Oil red O. Histochem. 97, 493-497.
15. Salminen, M. K., Tynkkyn, S., Rautelin, H., Saxelin, M., Vaara, M., Ruutu, P., Sarna, S., Valtonen, V., and Jarvinen, A. (2002) Lactobacillus bacteremia during a rapid increase in probiotic use of Lactobacillus rhamnosus GG in Finland. Clin. Infect. Dis. 35, 1155-1160.
16. Takemura, N., Okubo, T., and Sonoyama, K. (2010) Lactobacillus plantarum strain No. 14 reduces adipocyte size in mice fed high-fat diet. Exp. Biol. Med. 235, 849-856.
17. Tanida, M., Shen, J., Maeda, K., Horii, Y., Yamano, T., Fukushima, Y., and Nagai, K. (2008) High-fat diet-induced obesity is attenuated by probiotic strain Lactobacillus paracasei ST11 (NCC2461) in rats. Obes. Res. Clin. Pract. 2, 159-169.
18. Yanovski, S. Z. and Yanovski, J. A. (2002) Obesity. N. Engl. J. Med. 346, 591-602.