ABSTRACT

BACKGROUND AND PURPOSE: The repeatability of head CT interpretations may be studied in different contexts: in peer-review quality assurance interventions or in interobserver agreement studies. We assessed the agreement between double-blind reports of outpatient CT scans in a routine academic practice.

MATERIALS AND METHODS: Outpatient head CT scans (119 patients) were randomly selected to be read twice in a blinded fashion by 8 neuroradiologists practicing in an academic institution during 1 year. Nonstandardized reports were analyzed to extract 4 items (answer to the clinical question, major findings, incidental findings, recommendations for further investigations) from each report, to identify agreement or discrepancies (classified as class 2 [mentioned or not mentioned or contradictions between reports], class 1 [mentioned in both reports but diverging in location or severity], 0 [concordant], or not applicable), according to a standardized data-extraction form. Agreement regarding the presence or absence of clinically significant or incidental findings was studied with \( \kappa \) statistics.

RESULTS: The interobserver agreement regarding head CT studies with positive and negative results for clinically pertinent findings was 0.86 (0.77–0.95), but concordance was only 75.6% (67.2%–82.5%). Class 2 discrepancy was found in 15.1%; class 1 discrepancy, in 9.2% of cases. The \( \kappa \) value for reporting incidental findings was 0.59 (0.45–0.74), with class 2 discrepancy in 29.4% of cases. Most discrepancies did not impact the clinical management of patients.

CONCLUSIONS: Discrepancies in double-blind interpretations of head CT examinations were more common than reported in peer-review quality assurance programs.

ABBREVIATION: CHUM = Centre Hospitalier de l’Université de Montréal

The delivery of optimal radiology services may require continuous vigilance and perhaps quality assurance interventions.1–3 The content of these interventions may not be evident, however. In addition, the manner in which the error, discrepancy, and disagreement should be handled both in theory and in clinical practice is evolving.4 Discrepancies in peer-review approaches have been known for a long time.5–7 In 1959, Garland8 claimed that radiologists missed approximately 30% of tuberculosis cases in screening chest x-ray examinations.9 Garland’s report launched a series of investigations that continue today. However, there is no consensus on a standard method or protocol for evaluating errors and discrepancies in imaging reports, and rates published in the literature differ widely.1–3,10–14 Multiple variations in study parameters, including sampling sources, methods, imaging modalities, specialties, categories, interpreter training levels, and degrees of blinding, may have contributed to this wide spectrum.2,3,9

Recently, CT and MR imaging reports of the head, neck, and spine were re-read by staff neuroradiologists, and a 2% clinically significant discrepancy rate was found, an excellent result compared with the 3%–6% radiologic error rates published in general radiology practices.3,15,16

To anyone who has studied reliability or precision of diagnostic imaging tests, such levels of disagreement between interpretations may appear unbelievably low. Peer-review quality assurance “errors and discrepancies” and disagreements in reliability studies of imaging test interpretations may not measure the same things. Discrepancies in the reporting of imaging studies can thus be approached from at least 2 different perspectives.
From a quality assurance point of view, optimal radiology services require continuous quality assurance interventions. One report is the true right one, and discrepancies are errors that must be minimized. Performance can be measured; deviations and outliers can be identified, and appropriate measures can then be taken to improve performance.1-3

A different vocabulary is used when discrepancies are examined from a scientific point of view. In the typical absence of a criterion standard of “truth,” the uncertainty is a reality that must be admitted and taken into account when using imaging reports for clinical decisions. Reliability and agreement can be measured by using proper methods, including independent readings; and concordant or diverging verdicts can be tabulated and summarized, though imperfectly, by using marginal sums and appropriate statistical tools (such as $\kappa$ statistics). No test and certainly no imaging study requiring an element of interpretation will ever be perfectly repeatable.

Reconciliation between these 2 perspectives is desirable. The credibility of quality assurance programs disconnected from scientific methods is shaky. If only errors could be defined, perhaps as discrepancies beyond “normal discrepancies.” Unfortunately attempts to define an acceptable level of radiologic discrepancy are probably futile. Multiple variables are at play, and distinctions, even between acceptable discrepancy and negligence, may remain blurry.17

To our knowledge, reliability and agreement in the independent interpretation of head CT scans by expert neuroradiologists in a routine academic clinical practice have not been reported. In contrast to a peer-review approach, examining discrepancies after independent interpretations of clinical cases in everyday practice and looking for consensus on discrepant cases may provide a realistic and favorable framework for continuous quality improvement for each and all professionals, rather than the identification of specific deviant individuals. With this end in view, we studied the discrepancy in independent double readings of outpatient head CT scans in an academic practice. We hypothesized that our study would show a discrepancy rate in the range of $\geq 5\%$.

**MATERIALS AND METHODS**

The present article was written in accordance with the “Guidelines for Reporting Reliability and Agreement Studies” framework.18 A protocol was initially prepared, including a detailed data-collection form for each interpretation, prespecified plans for comparing pairs of reports, and statistical analyses. The protocol was discussed and accepted by all participating radiologists and by the head of the clinical neuroradiology service. Readers were informed that the identity of the participating radiologists would remain anonymous. As part of a pilot quality-improvement program, the necessity for obtaining informed consent was waived. Nonurgent outpatient head CT scans were interpreted on a double-blind basis within 2 consecutive days. The 2 reports were analyzed by 1 author (G.G.), who filled out the corresponding data-collection forms.

**Patients**

During 12 months (between July 2012 and July 2013), 119 outpatients (71 women; 48 men; mean age, 60.5 $\pm$ 15.4 years) with head CT requests from any outpatient clinic, including a neuro-oncology clinic ($n = 35, 29.4\%$), were randomly selected from the Centre Hospitalier de l’Université de Montréal (CHUM), Notre-Dame Hospital imaging library on a basis of 2–5 cases per week, on certain weeks only (cases could only be submitted to double reading when the first adjudicator was available to review reports in a timely fashion). No formal sample-size calculations were performed because this pilot project was considered exploratory, but a statistical consultant advised that $> 100$ patients were necessary to provide meaningful confidence intervals.

Cases were blindly and randomly selected and sent back, just as a “new case” would be interpreted within the workflow of the next consecutive day, for an independent re-reading. Cases that, by chance, happened to be read twice by the same reader were excluded from analyses ($n = 5$). CT scans from the emergency department or performed on hospitalized patients were excluded for methodologic and ethical considerations. The first report was automatically erased from the patient file when the second report became available but was saved in a separate file for this study. In this fashion, the second reader never had access to the first report, and head CTs were interpreted twice in a blinded fashion. All examinations were anonymously integrated in the daily workload, and both reports were dictated within 2 consecutive days. The second readings were made available to clinicians and became the permanent official report, unless flagged by G.G. When a finding was mentioned in the first but not in the second report, the second reader was asked to review or discuss the case with the first reader and amend the report if necessary. When the most inclusive or “safest” report was the second official one, no immediate revision was undertaken.

**Readers**

The readings were performed by 8 certified staff radiologists, fellowship-trained and specialized in neuroradiology with clinical experience ranging from 1 to 37 years (mean, 17.9 $\pm$ 7.3 years; median, 17 years). They worked independently, rotating each day within the same tertiary institution. They were aware that a study was ongoing but could not guess which patients were included (<1% of the patients being examined were part of the study).

**Data Extraction and Analyses**

Each report was analyzed by 1 author (G.G.) to extract the following 4 items: 1) response to the question raised by the requesting physician, 2) major/clinically pertinent findings (defined as related to the clinical question or requiring immediate management, such as recent ischemic lesions, tumor evolution or recurrence, acute sinusitis, and so forth), 3) incidental findings (defined as not being related to the clinical question and requiring no immediate management, such as chronic sinusitis, cerebral atrophy, old lacunar infarcts), 4) further investigations being proposed.

Pairs of reports were compared, and agreement was assessed regarding the following: 1) detection agreement or agreement on the presence or absence of an abnormal finding, sorting out positive-versus-negative test results; and 2) agreement on the description of the findings regarding the 4 previously mentioned items. For each of the first 3 items (clinical question, major clinically pertinent finding, and incidental finding), concordance was rated according to the content of the reports (and not according to the
provide proportions (with 95% confidence intervals). Class 2 discrepancies in reporting pertinent clinical findings from contrast-enhanced and nonenhanced CTs were compared by using the Fisher exact test. The Cohen $\kappa$ was calculated (with 95% confidence intervals) for the presence/absence of clinically pertinent findings and incidental findings. $\kappa$ values were interpreted according to Landis and Koch.19

RESULTS
There were 119 CT examinations; 53 (44.5%) were contrast-enhanced. A clinical question was formulated by the requesting physician in 98 cases (82.4%) and the reader responses were concordant in 88.1% of cases (95% CI, 80.6%–92.9%). The Cohen $\kappa$ for positive studies was 0.86 (0.77–0.95) and 0.59 (0.45–0.74) for the presence/absence of any clinically pertinent findings and for incidental findings, respectively.

Readings were in agreement for 75.6% (67.2%–82.5%), 65.5% (56.6%–73.5%), and 86.6% (79.2%–91.6%) of clinically pertinent findings, incidental findings, and recommendations for further investigations, respectively (Table 1). Rates were similar for the 2 subgroups (referred from the neuro-oncology clinic or from all other clinics).

Class 2 discrepancies in reporting clinically pertinent findings were found in 18 cases (15.1%). They are summarized in Table 2. Examples include the presence or absence of a tumor recurrence.

### Table 1: Agreement among double-blind reports

| Category                              | Concordant Readings (CI 95%) | Class 2 Discrepancies | Class 1 Discrepancies |
|---------------------------------------|------------------------------|-----------------------|-----------------------|
| Clinically pertinent findings         | 75.6% ($n = 90$) (67.2%–82.5%) | 15.1% ($n = 18$)      | 9.2% ($n = 11$)       |
| Incidental findings                   | 65.5% ($n = 78$) (56.6%–73.5%) | 29.4% ($n = 35$)      | 5% ($n = 6$)          |
| Answer to the clinician’s question    | 86.7% ($n = 85$) (78%–92.5%)  | 12.2% ($n = 12$)      | 1% ($n = 1$)          |
| Further investigations                | 86.6% ($n = 103$) (79.2%–91.6%) |                        |                       |

### Table 2: Class 2 discrepancies in clinically significant findings

| Age (yr) | Sex | NC or CE | Context                 | Discrepancies                  |
|----------|-----|----------|-------------------------|--------------------------------|
| 1        | 66  | F        | CE                      | F-up meningioma                | Progression vs stable          |
| 2        | 82  | F        | CE                      | F-up meningioma                | Progression vs stable          |
| 3        | 47  | M        | NC                      | F-up subdural hematoma         | Complete resolution vs partial regression |
| 4        | 60  | F        | NC                      | F-up hydrocephalus             | Stable vs progression          |
| 5        | 47  | F        | CE                      | F-up metastasis resection      | No recurrence vs recurrence suspicion |
| 6        | 77  | M        | CE                      | Lymphoma                       | Normal vs clival lesion         |
| 7        | 76  | M        | NC                      | F-up pituitary apoplexy        | Unchanged vs new hemorrhage     |
| 8        | 63  | M        | NC                      | F-up glaucoma                  | Optic nerve atrophy vs normal   |
| 9        | 66  | M        | NC                      | F-up subdural hematoma         | Persistence vs complete resolution |
| 10       | 64  | F        | NC                      | Post-op adenoma                | Residual mass vs normal         |
| 11       | 75  | F        | CE                      | Seizures                       | Schizencephaly vs no mention    |
| 12       | 74  | M        | NC                      | F-up subdural hematoma         | Rebleeding vs no rebleeding     |
| 13       | 58  | F        | CE                      | F-up metastases                | No recurrence vs recurrence     |
| 14       | 37  | F        | NC                      | F-up glioblastoma              | Hemiation vs no mention         |
| 15       | 62  | F        | CE                      | Breast CA                      | Choroidal mets suspicion vs no mention |
| 16       | 77  | F        | NC                      | Dementia                       | White matter disease vs no mention |
| 17       | 76  | M        | CE                      | Memory loss                    | Cerebral trophy vs no mention   |
| 18       | 42  | M        | CE                      | F-up glioblastoma              | Progression vs stable           |

Note: NC indicates noncontrast CT; CE, contrast-enhanced CT; CA, cancer; F-up, follow-up; Post-op, postoperative; mets, metastases.

**FIG 1.** Class 2 discrepancies in pertinent findings. **A,** A patient followed for recurrence of a resected breast cancer metastasis; one reader reported a recurrence (arrow), whereas the other observer did not mention this finding. **B,** A lytic lesion of the clivus (circle) was reported by one but not the other reader.

Clinical significance of the discordance) as the following: not applicable (in the absence of a clinical question or of any finding); 0, concordant; class 1 discordance, the same findings noted in both reports but in different locations or diverging in severity; class 2 discordance, the finding was not mentioned in one of the reports or opinions diverged on the evolution or recurrence of a lesion.

Subsequently, a senior author (J.R.) who had not previously read any of the examinations reviewed all reports and data-collection forms (unmasked).

**Statistical Analyses**

The concordance rates for responses to the clinical questions raised by the referring physicians, clinically pertinent findings, incidental findings, and proposed investigations were tabulated to
(n = 4, Fig 1A), the growth of a meningioma (n = 2), the evolution of chronic subdural hematomas (n = 2), and the presence of a lytic bone lesion (Fig 1B). These discrepancies were normally distributed between readers (n = 1, 3, 5, 7, 11, 5, 3, 1 for 36 discrepant reports). There was no significant difference between contrast-enhanced (n = 9 of 53) and nonenhanced studies (n = 9 of 66, P = .62).

Class 1 discrepancies in clinically pertinent findings were seen in the interpretation of 11 cases. Examples include the location of recent ischemic lesions (n = 2), tumor extensions (n = 2), or the disconnection of a ventricular shunt (n = 1).

Class 2 discrepancies in reporting incidental findings were seen in 35 cases. Examples include the presence or absence of white matter disease (n = 10), chronic sinusitis (n = 6), old strokes (n = 5), atrophy (n = 4), a nasal polyp (Fig 2A), or a pineal cyst (Fig 2B). Class 1 discrepancies in calling incidental findings were seen in 6 cases. Examples include the location of lacunar infarcts (n = 2) or the extension of chronic sinusitis (n = 2).

The senior author confirmed the discrepancies identified by the first adjudicator in all cases, except for 2 minor modifications in the categorization of incidental findings.

**DISCUSSION**

The salient findings of this study are the following: 1) The interobserver agreement regarding head CT outpatient studies with positive and negative findings measured by the Cohen κ was 0.86 (almost perfect), but 2) class 2 discrepancies in clinically pertinent findings were still found in 15% of cases, above the 5% level we expected from the 3%–6% discrepancy rates reported in previously published neuroradiologic peer-review studies.3,15,16

Different aspects of study design may explain our results. The definitions we used were somewhat arbitrary, though they were inspired from similar studies.3,16 Perhaps they were rigorously applied during data extraction. In the absence of standardization of reporting, some variability in the attribution of categoric verdicts to the content of reports of different styles is inevitable.

The levels of disagreement that we observed are not unheard of. Robinson et al20 investigated the concordance among 3 independent observers. Their study showed agreement in 51%, 61%, and 74% of abdominal, chest, and skeletal x-rays, respectively. They also assessed performance by calculating κ statistics of interobserver agreement. Weighted κ values between pairs of observers were higher with skeletal (0.76–0.77) than with chest (0.63–0.68) or abdominal (0.50–0.78) examinations. In a meta-analysis conducted by Wu et al,2 the global discrepancy rate was 7.7% (including a major discrepancy rate of 2.4%). The major discrepancy rate varied according to body region: It was lower for head (0.8%) and spine CT (0.7%) than chest (2.8%) and abdominal CT (2.6%). Blinding of the reference radiologist to the initial report was, however, associated with higher discrepancy rates compared with studies that lacked blinding: Not blinding the initial report yielded a much lower major discrepancy rate (2.0%; 95% CI, 1.4%–2.7%) than with blinding (12.1%; 95% CI, 4.4%–29.4%).

Hence, we believe that the main reason for the difference between our results and those of previous studies on discrepancies in head CT reports is that we used double-blind reporting. A possible explanation is that knowledge of the initial report may lead to a "satisfaction of search" error that reduces discrepancies.2

Most quality assurance studies published in the literature have assessed the discrepancy rates found through a peer-review approach.2,3,6,9,21 In that context, the opinion of the second observer is considered the criterion standard, and discrepancies are meant to be errors. However, in many cases, the second radiologist worked in a setting that differs from the normal clinical context. The second interpretation can be biased in many ways, by knowledge of the first interpretation, knowledge of the identity of the first interpreter, or even knowledge of the clinical evolution and outcome of the patient (ie, hindsight bias). Of course, the second interpreter is also aware that he or she is working within an audit process.3 Other studies that have used a double-blind method in a much larger number of patients have found lower discrepancy rates than those reported here.1 In this case, we suspect that the medical director or quality assurance radiologist knowingly working within an audit process may have biased, according to a principle of charity, the selective identification of the most serious discrepancies only. Our study being done in an anonymous, nonblaming context, with no intent to identify deviant individuals differs from many quality assurance studies and may explain why more frequent disagreements could be identified. Fortunately, most discrepancies had minimal clinical impact in terms of patient management or outcome.

Discrepancies in the present study do not necessarily mean errors. The use of clinical judgment, in assessing the impact of calling a suspicious-but-uncertain finding or the pertinence of reporting an incidental finding, for example, may lead to variations from one individual report to the other.

Some societies have proposed guidelines for peer-review

![FIG 2. Class 2 discrepancies in incidental findings. A, One reader reported a polypoid posterior nasal lesion (arrow), whereas the other observer did not mention this incidental finding. B, A pineal cyst (circle) was mentioned in only 1 of the 2 reports.](image-url)

AJNR Am J Neuroradiol 36:24–29 Jan 2015 www.ajnr.org 27
methods or for holding discrepancy meetings. Many questions remain unanswered, such as which method ought to be followed or which is most efficient. A survey of practices in the United Kingdom has recently shown wide variation in discrepancy meetings. Larson and Nance have emphasized that “peer review can either serve as a coach or a judge, but it cannot do both well.” There is a recent trend to shift away from the identification of deviant individuals to the overall improvement of systems and patient care, to move away from a blame culture to one of continuous feedback, learning from each other’s methods of working and reporting and from each other’s mistakes.

The work presented here was a pilot project. It was not designed to compete with commercially available peer-review systems. The process requires the active participation of a dedicated third physician (in addition to the 2 readers) and is time-consuming. In this particular case, the third physician was a senior resident, who considered the experience informative and enriching. Perhaps this pilot project could shed some light on how we should understand and design methods purported to improve patient care, including systematic quality assurance peer-review systems or discrepancy meetings. Neuroradiologists who participated in this project were initially surprised to see that discrepancies were common. They have started to exchange opinions on the pertinence of incidental findings, for example. The sharing of practices in an open setting of a multidisciplinary cancer center.

ACKNOWLEDGMENTS
The authors acknowledge the participation of the attending neuroradiologists of CHUM and the contribution of the PACS and imaging library staff.

REFERENCES
11. Agrawal A, Pandit M, Kalyanpur A.
12. Chung JH, Striegel RM, Chew AR, et al.
13. Goddard P, Leslie A, Jones A, et al.
14. Loevner LA, Sonners AI, Schulman BJ, et al.
15. Borgstede JP, Lewis RS, Bhargavan M, et al.
16. Siegel RL, Baram EM, Reuter SR, et al.

Enjoy the benefits of AJNR membership. Visit www.AJNR.org or call 877-226-3167 for more information.

Grants/Grants Pending
June 2015
Vol. 26, No. 6
AJNR: American Journal of Neuroradiology
1241 Washington Street, Stamford, Connecticut 06902
Telephone: 1-877-226-3167
Fax: 1-203-322-0619
E-mail: ajnr@ajnr.org
www.AJNR.org

Grants/Grants Pending
17. Berlin L. Radiologic errors and malpractice: a blurry distinction. AJR Am J Roentgenol 2007;189:517–22
18. Kottner J, Audige L, Brorson S, et al. Guidelines for Reporting Reliability and Agreement Studies (GRRAS) were proposed. J Clin Epidemiol 2011;64:96–106
19. Landis JR, Koch GG. The measurement of observer agreement for categorical data. Biometrics 1977;33:159–74
20. Robinson PJ, Wilson D, Coral A, et al. Variation between experienced observers in the interpretation of accident and emergency radiographs. Br J Radiol 1999;72:323–30
21. Herman PG, Gerson DE, Hessel SJ, et al. Disagreements in chest roentgen interpretation. Chest 1975;68:278–82
22. O’Keeffe MM, Piche S, Mason AC. The Canadian Association of Radiologists (CAR) Guide to Peer Review Systems. Approved September 10, 2011; Amended July 16, 2012. http://www.car.ca/uploads/standards%20guidelines/20120831_en_peer-review.pdf. Accessed November 21, 2013
23. The Royal College of Radiologists. Standards for radiology discrepancy meetings. London ETRCoR, 2007. https://www.rcr.ac.uk/docs/radiology/pdf/Stand_radiol_discrepancy.pdf. Accessed November 21, 2013
24. Prowse SJ, Pinkey B,Etherington R. Discrepancies in discrepancy meetings: results of the UK national discrepancy meeting survey. Clin Radiol 2014;69:18–22