Effectiveness of contact tracing in the control of infectious diseases: a systematic review

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Summary

Background Contact tracing is used for multiple infectious diseases, most recently for COVID-19, but data regarding its effectiveness in disease control are scarce. To address this knowledge gap and inform public health decision making for COVID-19, we systematically reviewed the existing literature to determine the effectiveness of contact tracing in the control of communicable illness.

Methods We searched PubMed, Embase, and the Cochrane Library from database inception up to Nov 22, 2021, for published studies evaluating associations between provider-initiated contact tracing for transmissible infectious diseases and one or three of interest: case detection rates among contacts or at the community level, overall forward transmission, or overall disease incidence. Clinical trials and observational studies were eligible, with no language or date restrictions. Reference lists of reviews were searched for additional studies. We excluded studies without a control group, using only mathematical modelling, not reporting a primary outcome of interest, or solely examining patient-initiated contact tracing. One reviewer applied eligibility criteria to each screened abstract and full-text article, and two reviewers independently extracted summary effect estimates and additional data from eligible studies. Only data reported in published manuscripts or supplemental material was extracted. Risk of bias for each included study was assessed with the Cochrane Risk of Bias 2 tool (randomised studies) or the Newcastle–Ottawa Scale (non-randomised studies).

Findings We identified 9050 unique citations, of which 47 studies met the inclusion criteria: six were focused on COVID-19, 20 on tuberculosis, eight on HIV, 12 on curable sexually transmitted infections (STIs), and one on measles. More than 2 million index patients were included across a variety of settings (both urban and rural areas and low-resource and high-resource settings). Of the 47 studies, 29 (61·7%) used observational designs, including all studies on COVID-19, and 18 (38·3%) were randomised controlled trials. 40 studies compared provider-initiated contact tracing with other interventions or evaluated expansions of provider-initiated contact tracing, and seven compared programmatic adaptations within provider-initiated contact tracing. 29 (72·5%) of the 40 studies evaluating the effect of provider-initiated contact tracing, including four (66·7%) of six COVID-19 studies, found contact tracing interventions were associated with improvements in at least one outcome of interest. 23 (46·9%) studies had low risk of bias, 22 (46·8%) studies had some risk of bias, and two (4·3%) studies (both randomised controlled trials on curable STIs) had high risk of bias.

Interpretation Provider-initiated contact tracing can be an effective public health tool. However, the ability of authorities to make informed choices about its deployment might be limited by heterogenous approaches to contact tracing in studies, a scarcity of quantitative evidence on its effectiveness, and absence of specificity of tracing parameters most important for disease control.

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Introduction

Contact tracing is defined by WHO as “the process of identifying, assessing, and managing people who have been exposed to a disease to prevent onward transmission”, which informs important follow-up actions such as diagnostic tests, post-exposure prophylaxis, and quarantine. Contact tracing has been used in the control of infections including smallpox, tuberculosis, HIV, other sexually transmitted infections (STIs), Ebola virus, and, now, COVID-19. The objectives of contact tracing include identifying potential new cases before they might infect others (with referrals to care and social supports when indicated), detecting clusters of cases before they expand, and improving overall understanding of disease dynamics (appendix p 2).

Although observers have highlighted apparent successes of contact tracing for COVID-19 in places including Taiwan, Vietnam, and US states such as Massachusetts, quantitative data are scarce on the strategy’s effectiveness for disease control in practice.
**Research in context**

**Evidence before this study**
Relatively few studies have attempted to quantify the effectiveness of contact tracing in the control of COVID-19 and other infectious diseases on the basis of experimental approaches or collection of observational data. Although many studies suggest contact tracing can interrupt COVID-19 spread, most studies base this conclusion either on mathematical modelling or narrative descriptions of programmes. We performed a preliminary search of PubMed, Embase, and the Cochrane Library for existing reviews evaluating contact tracing in the control of infectious diseases, using the search terms “contact tracing”, “contact investigation”, “contact examination”, “contact screen”, or “partner notification” or their permutations, and limiting the results to reviews and systematic reviews only that were published from database inception up to March 31, 2020, without language restrictions. This preliminary investigation revealed only a small number of existing systematic reviews assessing contact tracing for tuberculosis (six reviews) and HIV or other sexually transmitted infections (six reviews), which were unable to draw firm conclusions about the strategy’s overall effectiveness in population-level disease control. Additionally, these previous reviews each focused on only one type of infectious disease, revealing a missed opportunity to evaluate whether lessons from tracing programmes within specific contexts might help to inform best practices for the control of other illnesses, including novel diseases such as COVID-19. To address these gaps, we searched PubMed, Embase, and the Cochrane Library without date or language restrictions for published studies evaluating the effect of provider-initiated contact tracing in stopping infectious disease spread. Our search used the terms “contact tracing”, “contact investigation”, “contact examination”, “contact screen”, or “partner notification” or their permutations and included studies published from database inception up to Nov 22, 2021.

**Added value of this study**
This systematic review of contact tracing considers evidence of disease control effectiveness across multiple infectious diseases (COVID-19, tuberculosis, HIV, curable sexually transmitted infections, and measles). To our knowledge, this review is the first to synthesise empirical data for all forms of COVID-19 contact tracing and how they relate to disease control. We assessed data from more than 2 million index patients across eight infectious diseases and a variety of settings, including both urban and rural areas, and both low-resource and high-resource settings. 29 (72.5%) of 40 studies across all diseases, including four (66.7%) of six studies on COVID-19, found that provider-initiated contact tracing was associated with improvements in case detection, forward transmission, or disease incidence. However, the heterogeneity of approaches and the fact that most studies used observational designs limit our ability to understand how, where, and when to deploy this public health tool most effectively when disease control is the objective. Furthermore, the existing evidence base does not fully answer how parameters such as speed or completeness of contact tracing affect success, and how effectiveness might change in the event of case numbers exceeding health system capacity.

**Implications of all the available evidence**
The evidence suggests contact tracing to be an effective public health tool; however, more evidence is needed to understand how to optimise the effectiveness of contact tracing for infectious diseases control across a range of settings and contexts. Large-scale comparative studies of contact tracing that clearly specify approach and measure factors such as contact tracing speed and completeness would aid in decision making over resource allocation. In the interim, lessons from existing studies, such as the importance of ensuring privacy for contacts during tracing, and leveraging technology to increase tracing rates, can help to inform existing programmes for COVID-19 and other diseases.

Furthermore, contact tracing has implementation challenges in some settings; in a study of 74,185 patients with COVID-19 in the USA, 49,480 (66·7%) patients were not reached to elicit contacts or reported zero contacts.4 Previous systematic reviews focused on tracing for other infections, including tuberculosis and HIV,6 confirmed a high prevalence of these diseases among patients’ close contacts, yet remained unable to conclude the exact “contribution of contact tracing to reducing the burden of disease in the population”7. Improving our understanding of contact tracing’s effectiveness in disease control has substantial public health importance, particularly given ongoing SARS-CoV-2 transmission and the potential for future pandemics. Although the yield of contact tracing is likely to be influenced by the specific pathogen, population, and strategy involved, lessons from tracing programmes within specific contexts might help to inform best practices for disease control within and across disease categories (including for novel diseases such as COVID-19).

To address this knowledge gap and to inform public health decision making, we conducted a systematic review of published studies that used empirical data to evaluate the effectiveness of contact tracing to control the spread of infectious diseases.

**Methods**

**Search strategy and selection criteria**
We conducted a systematic review in accordance with the PRISMA guidelines.9 We searched PubMed, Embase, and the Cochrane Library for studies in the published literature starting from database inception up to Nov 22, 2021, with no date or language restrictions. Our search used the terms “contact tracing”, “contact
Articles

investigation”, “contact examination”, “contact screen”, or “partner notification” or their permutations. The full protocol and search strategy are provided in the appendix (pp 4, 21–23). We manually reviewed reference lists of reviews identified during the search for additional study articles to include.

We included clinical trials and observational studies evaluating the effect of contact tracing interventions specifically implemented by public health or health-care workers (ie, provider-initiated contact tracing, as previously defined;13–16 appendix p 5) on the control of transmissible infectious diseases. Transmissible infectious diseases were classed as all those transmitted through human-to-human contact. We included studies that had evaluated one of three outcomes of interest: case detection rates among contacts or at the community level (hypothesised to increase with contact tracing); overall forward transmission of disease as measured by the reproduction number (R; hypothesised to decrease), secondary attack rates (hypothesised to decrease), or similar measures (including cases and deaths prevented, reinfection rates among contacts, and treatment rates among contacts, where increased treatment rates were interpreted as a proxy for decreased forward transmission); and overall disease incidence (hypothesised to decrease). Other downstream disease outcomes, such as mortality, were considered but beyond the scope of this review. We included studies evaluating the effects of provider-initiated contact tracing compared with the absence of contact tracing or with patient-initiated contact tracing (appendix p 5). Studies evaluating expansions or programmatic adaptations of pre-existing provider-initiated tracing services and studies comparing provider-initiated tracing to tracing programmes without follow-up actions (such as instructions for contacts to present for testing or to avoid others for a specific period of time) were also included. We excluded studies that used mathematical modelling only; did not have a control group; did not report one of our outcomes of interest; or solely examined patient-initiated contact tracing.

We uploaded search results to Covidence software for deduplication, screening, and data extraction. After elimination of duplicate records, one reviewer (among ADH, JJ, AE, CXH, and AR) independently screened each abstract for full-text review and applied eligibility criteria to each full-text article; in cases of uncertainty, abstracts or full texts in question were reviewed by three reviewers and included or excluded by consensus.

Data analysis
Using a standardised form created for the review, two reviewers (among ADH, JJ, AE, CXH, and AR) independently extracted the following data from eligible studies: authors, infection studied, years of study, location and setting, study design, study population, sample size (numbers of index patients and contacts), details of contact tracing intervention, effect measured, and effect size. For the purposes of this study we considered summary effect estimates, with confidence intervals and p values when provided. Two reviewers (among the same group) then independently assessed risk of bias for each included study, using the Cochrane Risk of Bias 2 tool17 for randomised studies and the Newcastle–Ottawa Scale18 for non-randomised studies. Conflicts in data extraction or bias assessments were adjudicated by discussion between the two reviewers regarding each domain within the extraction form or bias tool to reach consensus, or by independent completion of the bias assessment by a third reviewer.

When possible, tracing and control approaches were categorised with consistent definitions (appendix p 5). Provider-initiated contact tracing included both provider referral (in which a trained health worker notifies contacts of their infection exposure and the need for evaluation or other actions) and contract referral (in which a trained health worker and index patient make an agreement that the health worker will not notify contacts until a prespecified time has passed during which time the index patient agrees to notify contacts and refer them for services themselves). Because tuberculosis is the most studied disease among all the diseases included, we additionally stratified tuberculosis studies by World Bank country classification (low income, lower-middle income, upper-middle income, or high income) and by burden of tuberculosis (high-burden or lower-burden settings, as defined in the appendix [p 6]).

We report on studies categorised by disease and type of contact tracing intervention (studies on provider-initiated contract tracing and those on programmatic adaptations within provider-initiated contact tracing), summarising the findings of each study and the risk of bias for each study. We expected that interventions and outcomes would be too heterogeneous for meta-analysis and thus did not plan to meta-analyse the data.

Role of the funding source
The funders of the study had no role in study design, data collection, data analysis, data interpretation, or writing of the report.

Results
We identified 14 254 records, including 5238 duplicates, from the database search (figure). 34 additional articles were identified by manually searching reference lists. We screened 9050 unique abstracts, yielding 361 full-text articles to be assessed for eligibility, of which 314 were excluded. Of the 47 included articles evaluating contact tracing interventions, six were focused on COVID-1919–25 (table 1), 20 on tuberculosis26–45 (table 2), eight on HIV46–53 (appendix pp 11–14), 12 on curable STIs (syphilis,54–56 gonorrhoea,57–60 trichomoniasis,61 chlamydia,62,63 and STI syndromes;57,58 appendix pp 15–19), and one on measles64 (appendix p 20). Included studies were conducted across
five continents (appendix p 3), with study years ranging from 1975 to 2020. The 47 studies comprised 18 (38.3%) randomised controlled trials (RCTs) and 29 (61.7%) observational studies. Of note, secondary attack rates might appear to be increased by contact tracing as a consequence of increased detection from tracing. A study from Portugal found that contact tracing led to a 2.26% (95% CI 7.1–18.9) or after (9.2% [7.8–10.8], p = 0.13). A separate matched-neighbours regression estimated that every percentage point increase in uptake of the same application across England and Wales between Oct 8 and Dec 31, 2020, led to a 2.26% (95% CI 1.50–3.00) reduction in COVID-19 cases. Another study from South Korea found widespread identification of individuals linked with clusters of COVID-19 cases was associated with a reduction in the time-varying reproduction number ($R_t$) from 1.3 (Jan 24 to March 8, 2020) to 0.6 (March 9 to May 2, 2020). The other two studies did not find evidence in support of COVID-19 contact tracing. An analysis of all COVID-19 cases from 130 countries (Jan to June, 2020) found only a weak relationship between contact tracing and $R_t$, specifically an association with higher $R_t$ values, which could not be well distinguished from other clustered interventions and was hypothesised as potentially reflecting increased case detection from tracing. A study from Portugal found no difference in COVID-19 secondary attack rates when tracing and quarantine measures were initiated either before laboratory confirmation of disease (12.1% [95% CI 7.1–18.9]) or after (9.2% [7.8–10.8], p = 0.13). Of note, secondary attack rates might appear to be increased by contact tracing as a consequence of increased detection of cases among contacts.

20 studies examined contact tracing for tuberculosis (table 2). These studies included more than 168 000 index patients across a diversity of settings, including urban and rural areas, high-burden and lower-burden areas (in terms of tuberculosis prevalence, incidence, or case notification rate), and all World Bank income levels (appendix pp 9–10). A variety of tracing approaches were described between and within studies, including household contact tracing (17 of 20 studies) and screenings of neighbours (three studies) and co-workers (one study;(table 2). The studies were primarily based in high-income countries, where additional social distancing measures were simultaneously in place. Four studies included a total of 1.9 million index patients, and the other two studies were population based. Fetzer and Graeber analysed a natural experiment enabled by a computer error in the UK in September, 2020, that delayed contact tracing for 15 841 index patients. By comparing areas more affected versus less affected by the computer error, they found that each late referral was associated with 18.6 additional COVID-19 cases and 0.24 additional deaths over 6 weeks. Two other UK-based studies also reported positive effects of contact tracing. On the Isle of Wight, initiation of traditional contact tracing and concurrent digital contact tracing (exposure notification via a smartphone application downloaded by more than 54 000 [38%] of 141 536 in the population) was associated with a reduction in the basic reproduction number ($R_0$) from 1.3 on May 5, 2020, to 0.5 on June 29, 2020. A separate matched-neighbours regression estimated that every percentage point increase in uptake of the same application across England and Wales between Oct 8 and Dec 31, 2020, led to a 2.26% (95% CI 1.50–3.00) reduction in COVID-19 cases. Another study from South Korea found widespread identification of individuals linked with clusters of COVID-19 cases was associated with a reduction in the time-varying reproduction number ($R_t$) from 1.3 (Jan 24 to March 8, 2020) to 0.6 (March 9 to May 2, 2020). The other two studies did not find evidence in support of COVID-19 contact tracing. An analysis of all COVID-19 cases from 130 countries (Jan to June, 2020) found only a weak relationship between contact tracing and $R_t$, specifically an association with higher $R_t$ values, which could not be well distinguished from other clustered interventions and was hypothesised as potentially reflecting increased case detection from tracing. A study from Portugal found no difference in COVID-19 secondary attack rates when tracing and quarantine measures were initiated either before laboratory confirmation of disease (12.1% [95% CI 7.1–18.9]) or after (9.2% [7.8–10.8], p = 0.13). Of note, secondary attack rates might appear to be increased by contact tracing as a consequence of increased detection of cases among contacts.

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| Country       | Study year | Design                          | Setting                                                                 | Sample                                                                                     | Intervention                                                                                     | Control                                                                                   | Outcome measured                                                                 | Results†                                                                                       | Risk of bias |
|--------------|------------|---------------------------------|------------------------------------------------------------------------|--------------------------------------------------------------------------------------------|-----------------------------------------------------------------------------------------------|-------------------------------------------------------------------------------------------|---------------------------------------------------------------------------------------------|-------------------------------------------------------------------------------------------------|--------------|
| Fetzer and Graeber (2021)<sup>20</sup> | UK         | 2020                            | Lower tier local authorities (districts, boroughs, and city councils) around the country (urban and rural) | 15 841 COVID-19 cases from Sept 25 to Oct 2, 2020, with contact tracing accidentally not performed on an estimated 48 000 contacts due to a computer error; areas were compared based on the degree they were affected by the error, which was seemingly random | Timely contact tracing for close recent contacts (not fully described) including instructions to self-quarantine; other social distancing policies in place not fully described | Delayed contact tracing for close recent contacts (not fully described), including delayed instructions to self-quarantine (computer error); other social distancing policies in place not fully described | Forward transmission: cases not prevented, deaths not prevented | Accidental failure to conduct timely contact tracing for 15 841 index patients (around 20% of index patients during the one-week period) was estimated to have led to an additional 126 836–185 188 additional COVID-19 cases during the 6 weeks after the error was discovered (22·5–32·8% of all cases during the period), as well as an additional 15·21–20·48 deaths over the 6 weeks (30·6–41·2% of all deaths during the period); ie, there were an estimated 18·6 additional COVID-19 cases per late referral and 0·24 additional deaths per late referral | Low |
| Kendall et al (2020)<sup>21</sup> | UK         | 2020                            | Isle of Wight compared with rest of the UK (urban and rural)           | All COVID-19 cases in the UK from March 28 to June 29, 2020 (exact number not specified) | Isle of Wight: traditional contact tracing (not fully described, starting May 5, 2020) and digital contact tracing with a Bluetooth-powered mobile application (starting May 7, 2020), with related advertising and community discussions (not fully described); pre-existing social distancing policies simultaneously in place (not fully described) | Rest of UK: initially no contact tracing, followed by traditional contact tracing only (not fully described, starting May 28, 2020); pre-existing social distancing policies simultaneously in place (not fully described) | Forward transmission: R<sub>t</sub> | After intervention initiation, the Isle of Wight had a decrease in R<sub>t</sub> (1·3 to 0·5) between May 5 and June 29, 2020, and also had a lower R<sub>t</sub> than the UK as a whole at the same timepoint (p<0.0001) | Low |
| Liu et al (2021)<sup>22</sup> | 130 nations | Country-level cohort study | 130 nations (urban and rural) | All COVID-19 cases in 130 nations from Jan 1 to June 22, 2020 | Contact tracing with intensities rated as comprehensive, limited, or none (not fully described; different countries likely to have employed different approaches); 12 other NPIs were also assessed and controlled for | No contact tracing or decreased intensity of contact tracing (not fully described) | Forward transmission: Rt | Contact tracing was associated with an increase in R<sub>t</sub> (exact effect size not specified) at 10 days; this increase might have reflected an increase in case detection, or other temporally associated NPIs in the model (eg, testing policies) | Low |
| Malheiro et al (2020)<sup>23</sup> | Portugal   | 2020                            | Eastern Porto (urban)                                                | All COVID-19 cases in eastern Porto from March 1 to April 30, 2020: 551 index patients, 1627 close contacts | Identification as a potential COVID-19 case through contact tracing (interview with patient or caregiver) or travel history, with mandatory quarantine prior to testing positive; social distancing policies (limiting movements and business activities) simultaneously in place starting March 22, 2020 | Identification as a COVID-19 case through testing, with no mandatory quarantine prior to testing positive; social distancing policies (limiting movements and business activities) simultaneously in place starting March 22, 2020 | Forward transmission: secondary attack rate | There was no significant difference in the secondary attack rate for index cases between the intervention group (16 of 132, 12 1% [95% CI 7·1–18·9]) and the control group (138 of 1495, 9·2% [7·8–10·8], p=0·13), including when stratifying by the presence or absence of social distancing policies (p=0·72) | Some |

(Table 1 continues on next page)
Table 1: Summary of included studies of contact tracing* for COVID-19

| Country       | Study year | Design                        | Setting                        | Sample                                                                 | Intervention                                                                 | Control                                                      | Outcome measured                      | Results†                            | Risk of bias |
|---------------|------------|-------------------------------|--------------------------------|------------------------------------------------------------------------|-----------------------------------------------------------------------------|----------------------------------------------------------------|--------------------------------------|-------------------------------------|--------------|
| South Korea   | 2020       | Pre-post design                | Seoul (urban)                  | All COVID-19 cases in Seoul from Jan 24 to May 2, 2020: 637 index patients, 16,176 contacts | Post-period (March 9 to May 2, 2020): widespread testing of contacts associated with clusters of COVID-19 cases identified via interviews, global positioning system data, credit card histories, drug utilisation review (not fully described), closed-circuit television records, and possibly other methods and individuals with COVID-19 symptoms; all positive-testing individuals moved to health-care settings for quarantine; other social distancing policies in place not fully described | Pre-period (Jan 24 to March 8, 2020): widespread testing of individuals with COVID-19 symptoms only; all positive-testing individuals moved to health-care settings for quarantine; other social distancing policies in place not fully described | Forward transmission: Rt decrease in Rt across Seoul (1.3 during the pre-period, 0.6 during the post-period) | Some       |
| UK            | 2020       | Retrospective cohort study     | All of England and Wales (both urban and rural) | All COVID-19 cases in England and Wales from Oct 8 to Dec 31, 2020: 18,920,000 index patients (325,900 deaths) | Digital contact tracing with a Bluetooth-powered mobile application, with simultaneous manual contact tracing and additional restrictive and lockdown measures (not fully described) in place | No digital contact tracing, but with simultaneous manual contact tracing and additional restrictive and lockdown measures (not fully described) in place | Forward transmission: cases prevented, deaths prevented | With matched-neighbours regression, use of the contact tracing application by 16.5 million (27.8% of 59.4 million people in England and Wales) was estimated to prevent 594,000 COVID-19 cases (95% CI 317,000–914,000) and 8,700 deaths (95% CI 4,700–13,900) between Oct 8 and Dec 31, 2020; each percentage point increase in application usage was associated with a 2.26% (95% CI 1.50–3.00) reduction in cases for the total period (1.09% [0.04–2.14] reduction from Oct 8 to Nov 6 and 2.66% [1.75–3.56] reduction from Nov 7 to Dec 31 following application improvements released on Oct 29 that were not fully described) | Low       |

*R0 = reproduction number. Rt = time-varying reproduction number. NPI = non-pharmaceutical intervention. *Provider-initiated contact tracing for all studies listed. †Significance or non-significance of result is not stated when it was not specified in the study.
| Country      | Study years | Design                | Setting                                                                 | Sample            | Intervention                                                                                                                                   | Control                                                                 | Outcome measured                                                                 | Results*                                                                 | Risk of bias |
|-------------|-------------|-----------------------|-------------------------------------------------------------------------|-------------------|-----------------------------------------------------------------------------------------------------------------------------------------------|--------------------------------------------------------------------------|-----------------------------------------------------------------------------|--------------------------------------------------------------------------|--------------|
| Ayles et al (2013) | Zambia, South Africa | 2010 Cluster RCT (two-by-two factorial design) | 16 communities in Zambia and eight communities in South Africa (urban and rural) | 64,463 individuals screened for tuberculosis | Three intervention groups: (1) complex household intervention including household contact tracing for tuberculosis, as well as HIV testing for household contacts and appropriate linkages to care; (2) complex community intervention including community mobilisation (eg, pamphlets, megaphone announcements) and establishment of additional and expedited sputum collection sites; (3) both the complex household intervention and complex community intervention | Passive case finding for all groups (intervention and control) also received improved clinical-based tuberculosis and HIV services (eg, diagnostics) | Overall active tuberculosis prevalence (among adults), and overall latent tuberculosis incidence (among children) | When comparing communities in the study groups with and without the complex/household intervention (which included household contact tracing), there was no significant difference in the prevalence of active tuberculosis among adults 4 years after the start of the intervention (adjusted prevalence ratio 0.82 [95% CI 0.64–1.04], p=0.095, permutation test p=0.10), or in the incidence of latent tuberculosis among children (adjusted IR 0.45 [95% CI 0.20–1.05], p=0.063, permutation test p=0.10) | Low          |
| Becerra et al (2005) | Peru | 1996–97 Prospective cohort study | Low-income neighbourhood in Lima (urban) | 208 index patients, 1,994 household contacts, 2,253 neighbours | Household contact tracing and screening of neighbours | Passive case finding of index patients told to instruct household contacts to present to clinic if the contacts developed tuberculosis-like symptoms | Case detection among contacts | A higher tuberculosis prevalence among households in the study group (eight [0.7%] of 1,094) than by passive case finding (two [0.1%] of 1,094), the tuberculosis prevalence among neighbours of index cases was 0.1% for both groups (three of 2253 and two of 2253, respectively) | During the 5-year study period, the incidence of tuberculosis in neighbourhoods receiving DOTS with household contact tracing (DOTS-A) decreased by 10.0% (from 339 to 305 cases per 100,000 residents), while the incidence in neighbourhoods receiving conventional DOTS increased by 5.3% (from 340 to 358 cases per 100,000 residents), resulting in a 14.8% lower incidence in DOTS-A neighbourhoods compared with conventional DOTS neighbourhoods (p=0.04) | Some         |
| Cavalcante et al (2010) | Brazil | 2000–04 Cluster RCT | Eight neighbourhoods in Rio de Janeiro, pair-matched according to tuberculosis incidence (urban) | 650 index patients, 2,147 household contacts | Enhanced DOTS with household contact tracing (DOTS-A), including treatment for contacts found to have active tuberculosis and prophylaxis for those found to have latent tuberculosis infection | Conventional DOTS with patient-initiated contact tracing | Overall disease incidence | During the 5-year study period, the incidence of tuberculosis in neighbourhoods receiving DOTS with household contact tracing (DOTS-A) decreased by 10.0% (from 339 to 305 cases per 100,000 residents), while the incidence in neighbourhoods receiving conventional DOTS increased by 5.3% (from 340 to 358 cases per 100,000 residents), resulting in a 14.8% lower incidence in DOTS-A neighbourhoods compared with conventional DOTS neighbourhoods (p=0.04) | Low          |
| Dongo et al (2021) | Uganda | 2014–16 Pre-post design | Two districts: Kabarole (rural) and Wakiso (peri-urban) | All detected tuberculosis cases in Kabarole and Wakiso (exact number not specified) | Post-period (July, 2015 to Dec, 2016): decentralisation of child tuberculosis services, which included household contact tracing after training of 178 CHWs, and health system improvements including repairs to laboratory equipment and training for laboratory staff | Pre-period (Jan, 2014, to June, 2015): child tuberculosis services prior to decentralisation, including before household contact tracing and before other health system improvements | Case detection at community level | There was a 138.7% increase in tuberculosis cases in children (age 0–14 years) detected during the post-period (647 cases), compared with the pre-period (271 cases); there was also a 31.7% increase in tuberculosis cases detected among adolescents and adults aged ≥15 years during the post-period (3693 cases) compared with the pre-period (2805 cases) | Some         |

(Table 2 continues on next page)
### Country Study years Design Setting Sample Intervention Control Outcome measured Results Risk of bias

| Country | Study years | Design | Setting | Sample | Intervention | Control | Outcome measured | Results | Risk of bias |
|---------|-------------|--------|---------|--------|--------------|---------|-----------------|---------|--------------|
| Nigeria | 2017–19     | Prospective cohort with historical comparison | Three states in southern Nigeria: Akwa Ibom State (intervention state), Cross River State (intervention state), and Rivers State (control state; none specified as urban or rural) | 509,768 individuals screened for tuberculosis (both household contacts of index patients and individuals screened in other places) | Several simultaneous interventions including household contact tracing; screening of individuals in other homes and tents (H2H/T2T); screening of individuals in public areas (community screening); and promotion of the interventions via print media and public talks | Baseline tuberculosis services (not fully described) in the control state during the study period, and in the intervention states prior to the study period | Case detection at community level | Intervention states saw a 112·9% increase in all forms of tuberculosis detected versus baseline (1258 cases in 2018–19, versus 591 cases in 2017–18), or a 138·3% increase compared with the expected trend; this increase in all forms of tuberculosis detected was greater than the increase observed in the control state (410 cases in 2018–19), versus 204 cases in 2017–18, or a 101·0% increase), or a 49·1% increase compared with the expected trend; household contact tracing resulted in a greater proportion of screened individuals diagnosed with presumptive tuberculosis (1350 [11·5%] of 11,700) compared with H2H/T2T (5786 [2·2%] of 269,069) and community outreach (4536 [2·2%] of 209,177) | Some |
| Ethiopia | 2011–14 | Cross-sectional study | Six zones in Oromia and Amhara regions (rural) | 47,021 index patients, 272,441 close contacts | Retrospective tracing of close contacts for all tuberculosis cases in the previous 3 years, which included household contact tracing | Routine case detection strategy (not fully described) | Case detection among contacts | A higher tuberculosis prevalence among close contacts was detected by retrospective contact tracing (768 per 100,000) compared with the control strategy (baseline case notification rate reported to be more than 130 per 100,000 for these zones) | Low |
| Nepal | 2017–18 (historical cohort study 2014–17) | Prospective cohort study | Eight districts (not specified if urban or rural) | 54,239 individuals screened for tuberculosis | Three-part active case finding intervention: (1) contact tracing of close contacts (not fully described); (2) facility-based screening; (3) tuberculosis cases in various underserved communities where attendees were screened for symptoms | Passive case finding and limited household contact tracing (not fully described) | Case detection at community level | The intervention was associated with a 22·3% increase in detected tuberculosis cases in intervention districts compared with control districts, and was also associated with a 29·0% increase in detected tuberculosis cases within intervention districts compared with the previously expected trend in intervention districts | Some |
| South Africa | 2017–19 | Cluster RCT | Catchment area of 56 primary care clinics in Limpopo province (rural) | 3665 index patients, 1677 contacts | Household contact tracing, or contacts invited for screenings at clinics with monetary incentive | Facility-based screening of all patients | Forward transmission: treatment initiation rates | There was no significant difference in tuberculosis treatment initiation rates between the facility-based screening group and household contact tracing group (treatment initiation rate ratio 1·04 [95% CI 0·83–1·30], p=0·73) | Low |
| Mexico | 1990–2010† | Longitudinal population-based study | Whole country (urban and rural) | 72,398 index patients with all forms of tuberculosis in Mexico (2007–10) | Contact investigation (not fully described) | Modelled the control group based on average number of contacts examined per reported tuberculosis case | Overall disease incidence | There was a negative correlation between the mean number of contacts examined per case and the incidence of tuberculosis (all forms; $r^2$=0·47, p=0·010), with a 33·0% reduction in (all forms of tuberculosis) incidence for each unit increase in the mean number of contacts examined | Some |

(Continued from previous page)

* (Table 2 continues on next page)
| Country          | Study years | Design                     | Setting                                      | Sample                        | Intervention                                      | Control                | Outcome measured  | Results* | Risk of bias |
|------------------|-------------|----------------------------|----------------------------------------------|-------------------------------|--------------------------------------------------|------------------------|--------------------|----------|--------------|
| Khatana et al    | 2014-15     | Quasi-RCT†                 | Two communities in Kashmir (rural)          | 282 index patients, 1191 household contacts | Household contact tracing                       | Passive case finding   | Case detection among contacts | A higher tuberculosis prevalence among household contacts was detected by household contact tracing (27 [4.5%] of 598) compared with passive case finding (seven [1.2%] of 539; OR 3.97 [95% CI 1.73-9.11], p=0.001) | Some |
| Kliner et al     | 2011-12     | Prospective cohort study   | Catchment area of a regional hospital (rural) | 122 index patients, 658 household contacts | Three sequential interventions for household contacts: (1) phone call reminders for contacts asking them to present for screening, followed by household contact tracing if contact does not present (November, 2011-February, 2012); (2) screening of contacts by phone, followed by household contact tracing if contact is not reachable by phone (March-July, 2012); (3) household contact tracing by contact referral (7 days) for high-risk contacts (>5 years old, index patient with multidrug-resistant tuberculosis, or index patient with a high-grade [3+ grade] sputum smear), or a second invitation letter given to contacts not deemed high-risk (August-October, 2012) | Patient-initiated contact tracing, and screening of any contacts who accompanied index patients to clinic | Case detection among contacts | A greater number of tuberculosis cases was detected by patient-initiated contact tracing and screening of contacts who accompanied index patients to clinic (four [2.5%] of 157 screened contacts), compared with any of the intervention methods (0 of 137 contacts in the phone call intervention, 0 of 226 contacts in the phone screening intervention, 0 of 107 high-risk contacts in the household contact tracing by contract referral intervention, and 0 of 31 contacts who received a second invitation letter) | Low |
| Mandalakas et al | 2013-15     | Prospective cohort study   | Seven basic management units (urban and rural) | 3258 index patients, 12,175 household contacts | Household contact tracing | No household contact tracing before the intervention (not fully specified) | Case detection at community level | The first year of the household contact tracing intervention was associated with a 32% increase in the number of children diagnosed with bacteriologically confirmed tuberculosis despite a concurrent decrease in the national tuberculosis case notification rate | Low |

*(Table 2 continues on next page)*
| Country       | Study years | Design                  | Setting                                          | Sample                                      | Intervention                                                                 | Control                                      | Outcome measured                                      | Results*                                                                 | Risk of bias |
|---------------|-------------|-------------------------|-------------------------------------------------|---------------------------------------------|--------------------------------------------------------------------------------|---------------------------------------------|----------------------------------------------------------|--------------------------------------------------------------------------|--------------|
| Cambodia      | 2012–14     | Cluster RCT (quasi-experimental) | 30 districts with high rates of poverty (not specified if urban or rural) | 25 668 index patients                      | Household contact tracing and screening of neighbours                          | Passive case finding                      | Case detection at community level                            | Household contact tracing was associated with an increase in all forms of tuberculosis detected compared with baseline: 9.8% increase during year 1, when 30 of 30 districts were conducting the intervention; 23.4% increase during year 2, when 30 of 30 districts were conducting the intervention; 6 quarters after year 1, there was a decrease in all forms of tuberculosis detected (−218%) compared with the previously expected trend | Low          |
| Afghanistan   | 2011–12     | Prospective cohort study | Six provinces (not specified if urban or rural) | 2 022 127 health facility attendees, household contacts of patients with tuberculosis and internally displaced people | Household contact tracing                                                      | Facility-based screening of all patients, and screening all people in internally displaced person camps | Case detection among contacts                             | A higher prevalence of sputum smear-positive tuberculosis was detected by household contact tracing (268 [1.6%] of 16 645 household contacts) than by other strategies (4125 [0.2%] of 16 992 277 people presenting to health facilities with any concern, 653 [0.2%] of 306 265 people in internally displaced person camps) | Some         |
| Peru          | 2012–14     | Cluster RCT (stepped-wedge design) | Densely populated district in Lima (urban) | 3 222 index patients, 12 566 household contacts | Household contact tracing                                                      | Passive case finding, index patients told to instruct household contacts to present to clinic if these contacts developed tuberculosis-like symptoms | Case detection among contacts                             | A higher incidence of tuberculosis among household contacts was detected by household contact tracing than by passive case finding (IRR 1.51 [95% CI 1.21–1.88], p=0.0001 for pulmonary tuberculosis; IRR 1.48 [1.19–1.84], p=0.0001 for all forms of tuberculosis) | Some         |
| USA           | 2012        | Observational, with hypothetical control | All 50 states and Puerto Rico (urban and rural) | 9945 index patients, 105 100 contacts (not fully described) | Contact investigation (not fully described) with treatment of detected latent tuberculosis infection | No contact investigation or treatment of detected latent tuberculosis infection | Forward transmission: cases presented                    | Contact investigation and treatment of latent tuberculosis infection was estimated to prevent 128 cases (95% CI 64–252) of tuberculosis over 5 years | Some         |
| Malawi        | 2001–02     | Pre-post design          | One district (rural)                            | 189 index patients, 985 household contacts    | Post-period (2002): passive case finding (household visit performed only to record information on household contacts) | Case detection among contacts               | A higher prevalence of tuberculosis among household contacts was observed during the post-period (eight [1.74%] of 461) compared with the pre-period (one [0.19%] of 524, p=0.01) | Some         |

(Table 2 continues on next page)
17 studies evaluated the effect of provider-initiated contact tracing for tuberculosis compared with a control intervention. Four of six RCTs and quasi-RCTs comparing contact tracing for tuberculosis to passive case finding or facility-based screening found an association between contact tracing and an increase in case detection\(^\text{30,31,41,43}\) or decrease in incidence (when tracing was coupled with treatment for contacts with active disease and prophylaxis for contacts with latent infection)\(^\text{29}\), whereas, the other two RCTs did not observe a positive effect on outcomes of interest\(^\text{30,31}\) (table 2). Among 11 observational studies comparing tracing programmes with control interventions, eight (72.7%) found an increase in tuberculosis case detection\(^\text{27,30,32,34,35,40,42,45}\), one (9.1%) identified an increase in cases averted\(^\text{33}\), one (9.1%) found a decrease in incidence\(^\text{36}\), and one (9.1%) found fewer cases detected with provider-initiated tracing than with patient-initiated tracing\(^\text{37}\) (table 2). Among these 17 studies overall, tracing for tuberculosis led to improved outcomes in low-burden settings (three of three)\(^\text{27,38,41}\) and high-burden settings (eleven of 14 studies)\(^\text{26–28,32,36–40,43,44}\) and in low-income countries (six of six)\(^\text{30,34,35,41,42,45}\) and countries in other income categories (eight of 11 studies)\(^\text{26–28,32,36–40,43,44}\) (table 2, appendix pp 9–10). The largest study in scale examined annual data from 2007 to 2010 for reported tuberculosis cases across Mexico, finding that each unit increase in the mean number of contacts traced per patient associated with a 33.0% decrease in national incidence (\(r^2=–0.47,\ p=0.010\))\(^\text{37}\).

Three studies compared programmatic adaptations within provider-initiated contact tracing for tuberculosis (table 2). A cluster RCT in Uganda found no difference in the odds of tuberculosis diagnosis among household contacts when sputum was collected at home (intervention group) versus when contacts were referred to clinics for sputum collection (odds ratio 1.34 [95% CI 0.42–4.24], \(p=0.62\), with the authors noting that only 35 (38.5%) of 91 eligible contacts in the intervention group could produce sputum samples at home\(^\text{29}\). Two observational studies evaluated expansions to household contact tracing alone. Three studies compared programmatic adaptations within provider-initiated contact tracing for tuberculosis (table 2). A cluster RCT in Uganda found no difference in the odds of tuberculosis diagnosis among household contacts when sputum was collected at home (intervention group) versus when contacts were referred to clinics for sputum collection (odds ratio 1.34 [95% CI 0.42–4.24], \(p=0.62\), with the authors noting that only 35 (38.5%) of 91 eligible contacts in the intervention group could produce sputum samples at home\(^\text{29}\). Two observational studies evaluated expansions to household contact tracing alone. Three studies compared programmatic adaptations within provider-initiated contact tracing for tuberculosis (table 2). A cluster RCT in Uganda found no difference in the odds of tuberculosis diagnosis among household contacts when sputum was collected at home (intervention group) versus when contacts were referred to clinics for sputum collection (odds ratio 1.34 [95% CI 0.42–4.24], \(p=0.62\), with the authors noting that only 35 (38.5%) of 91 eligible contacts in the intervention group could produce sputum samples at home\(^\text{29}\). Two observational studies evaluated expansions to household contact tracing alone. Two observational studies evaluated expansions to household contact tracing alone. Two observational studies evaluated expansions to household contact tracing alone.

Eight studies examined the effect of contact tracing programmes for HIV, enrolling a total of 15 439 index patients (appendix pp 11–14).\(^\text{5,6,8}\) Five RCTs, two from Malawi\(^\text{8,9,10}\) and one each from China,\(^\text{10}\) Kenya,\(^\text{9}\) and the USA,\(^\text{10}\) found that provider-initiated contact tracing was associated with higher HIV case detection rates compared with patient-initiated contact tracing. However, the study in China could not separate the effect of provider-initiated tracing from patient-initiated tracing with self-testing kits.\(^\text{10}\) Among the observational studies,

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a pre–post study from Botswana found that offering provider-initiated contact tracing and other services to index patients did not increase the number of partners diagnosed with HIV per index patient compared with patient-initiated tracing (0.14 vs 0.13, p=0.50).69 Two cohort studies from the USA evaluated adaptations to HIV contact tracing programmes. Malave and colleagues showed that contact tracing by disease intervention specialists (DIS) versus community clinicians led to similar rates of new HIV diagnoses among tested partners (20 [27·0%] of 74 partners tested and 10 [22·2%] of 45, p=0.56), although DIS intervention elicited significantly more partners per index patient (0.87 vs 0.22, p<0.01).62 Udeagu and colleagues3 reported a reduced HIV case detection rate with provider-initiated tracing by mobile phone text message (five [1·5%] cases among 325 contacts) or email (three [1·1%] among 267), compared with traditional methods (telephone or in-person; 106 [5·3%] among 2009).

12 studies examined provider-initiated contact tracing for curable STIs, which included 2383 index patients for syphilis,64–66 45 228 for gonorrhoea,57–60 484 for trichomoniasis,61 862 for chlamydia,62,63 and 1446 for non-specified STIs64,65 (appendix pp 15–19). Five RCTs or quasi-RCTs compared provider-initiated contact tracing with control interventions. Two RCTs, one on chlamydia in the USA66 and one on syndromic STIs in Zambia,67 found that provider-initiated tracing for male patients led to an increase in the number of treated partners (as a proxy for forward transmission); however, the Zambian study did not specify how many participants chose provider-initiated contact tracing versus an alternative referral slip (patient-initiated) intervention within the overall intervention group,67 and the US study did not assess how many partners in the control group received care elsewhere (and was therefore rated at high risk of bias).66 Two other RCTs and one quasi-RCT did not find evidence in favour of provider-initiated contact tracing for STIs,66,64,65 although one study noted a potentially underpowered analysis due to low trichomoniasis reinfection rates.62 Among the five observational studies comparing provider-initiated contact tracing with other interventions, four on gonorrhoea or syphilis found that provider-initiated tracing was associated with improvements in one or more of our outcomes of interest.57,58,61 In the remaining observational study on chlamydia, findings significantly fewer partners completed treatment after provider-initiated tracing by DIS compared with a patient-initiated approach.68 In addition, two studies examined the effect of programmatic adaptations within provider-initiated contact tracing interventions: an RCT detected no difference in syphilis case detection between immediate provider referral compared with control referral (but was assessed to have high risk of bias due to absence of randomisation concealment and deviations in administered interventions; appendix p 8),68 and an observational study found the proportion of syphilis patients with at least one partner treated increased from 26% to 28% (8% relative increase) when tracing by email was added to traditional in-person or telephone contact.57

The single observational study for measles took place in the USA during a measles outbreak in 2017. It assessed contact tracing plus recommendation to isolate if symptoms develop, compared with contact tracing plus exclusion, in which measles-exposed individuals were recommended to be temporarily excluded from high-risk settings and to avoid public places, public transportation, and home visitors (appendix p 11). The Rt for non-excluded individuals was 1·61 (95% CI 1·00–2·69) compared with 0·38 (0·20–0·73) for excluded individuals, for a relative transmissibility (unadjusted ratio of Rt values) of 4·2 (95% CI 1·9–9·6).66

Discussion

We identified 47 empirical studies that either evaluated the effectiveness of provider-initiated contact tracing or compared programmatic adaptations within contact tracing interventions,57–66 including six studies focused on COVID-19. Among the studies were a wide range of tracing strategies (including phone calls,57,58,61,63 text messaging,57,62,63 and household contact tracing57,58,61,63), and personnel involved with the contact tracing (ranging from public health specialists to volunteers), highlighting the diversity of programmatic approaches often taken.

Overall, provider-initiated contact tracing was generally associated with improved control of communicable illnesses: 29 (72·5%) of 40 studies evaluating the effect of provider-initiated contact tracing, including four (66·7%) of six observational studies on COVID-19 and 11 (68·8%) of 16 RCTs and quasi-RCTs on other diseases, reported increased case detection, decreased forward transmission, or decreased disease incidence compared with control interventions. Tracing was associated with improved outcomes across pathogens with diverse transmission dynamics and across a variety of settings. However, our ability to use these observations to form precise estimates on the effectiveness of contact tracing in disease control is limited largely by the heterogeneity of approaches and also by a relative scarcity of RCT data. Most studies (29 [61·7%] of 47) used observational designs with inherent vulnerability to confounding and other biases.

Among studies of contact tracing for COVID-19, four of six studies reported notable reductions in cases,57,62,63 Rr21, or Rt2 after provider-initiated contact tracing efforts, suggesting that such programmes can have effectiveness in mitigating disease spread. However, these studies were almost exclusively in high-resource settings57,62,63,25 and used observational designs with differing programme approaches,25 limiting generalisability. Most studies took place during a dynamic public health situation; as such, fluctuations in COVID-19 transmission
and other externalities might have reduced accuracy in estimations of tracing effectiveness. These six studies do not capture how tracing speed or completeness influence success, and such parameters are likely to be important for disease control when pre-symptomatic onward transmission occurs. The studies also do not quantify how tracing feasibility and effectiveness might change as cases exceed a contact tracing workforce or laboratory testing capacity.

Most studies of provider-initiated contact tracing for diseases other than COVID-19 similarly reported improvements in our outcomes of interest, suggesting some benefit in disease control. The repeatedly reported effectiveness of tracing interventions for tuberculosis, including in low-resource, high-burden settings, was notable. However, these studies similarly examined many distinct tracing interventions and largely relied on observational methods, once again preventing a summary effectiveness measure.

Of note, two studies found that provider-initiated contact tracing resulted in lower case detection or fewer partners completing treatment than patient-initiated tracing. Although inconsistent with most other identified studies, these findings suggest that in particular contexts, well designed, patient-initiated methods might be preferable to provider-initiated approaches. These results might alternatively reflect difficulties in connecting provider-traced contacts to follow-up services in these studies.

As a result, although contact tracing is shown to be an effective public health tool, additional evidence is needed to inform how, where, and when to deploy this tool most effectively for maximum effect in disease control, including against COVID-19. When feasible, further comparative evaluations, such as randomised, pragmatic trials within existing tracing systems, should be undertaken to evaluate different tracing interventions in real life and identify potential adaptations for unique infections and contexts.

In the interim, specific lessons learned from these studies, including those in which provider-initiated contact tracing was not shown to reduce disease spread or appeared to worsen outcomes, can inform elements of tracing programmes for COVID-19 and other diseases. Davis and colleagues concluded that low rates of home-based contact tracing for tuberculosis might have related to limited experience among community health workers in collecting sputum and a hesitation among contacts to expectorate in public, highlighting the importance of sufficient training for tracers and privacy for contacts. Two studies (Udeagu et al and Ehlman et al) showed that contact tracing via email or text message had less success than traditional contract tracing, but nevertheless increased case detection or the number of index patients with at least one partner receiving treatment when added to traditional contact tracing, highlighting how technology might increase tracing impact. Notably, two UK studies found widespread use of a smartphone-based proximity detection and exposure notification application was associated with substantial reductions in COVID-19 cases and \( R_0 \), although these studies were unable to fully isolate the effects of digital contact tracing from simultaneous traditional contact tracing and other interventions.

Strengths of this review include its consideration of multiple geographic regions and socioeconomic contexts across the included studies. Limitations include the predominance of observational studies, incomplete reporting in some studies, and the possibility of publication bias. We aimed to better understand how contact tracing would apply to the COVID-19 pandemic, but the unique characteristics of COVID-19, including frequent presymptomatic transmission and a relatively short serial interval, limit the generalisability of evidence from other diseases. Our study focused on disease control and did not evaluate other important outcomes of contact tracing such as how it helps understanding of disease dynamics or its ability to serve as an entry point to the health system for those at risk of disease. Further research should evaluate which patient groups (such as individuals experiencing homelessness or marginalised or minoritised communities) might differentially be excluded from or could be engaged by tracing programmes; how tracing might be affected by local values or customs, or by the accuracy and privacy of digital tools, and how the cost and cost-effectiveness of tracing programmes compare with alternative interventions.

Our review confirms a premise for contact tracing as an effective public health tool for infectious disease control, however we found little empirical data to inform best practices for contact tracing, revealing an opportunity to further evaluate this common practice relative to other public health strategies. As health agencies continue their responses against COVID-19 while responding to other ongoing health threats and planning for future pandemics, addressing these knowledge gaps through additional studies or analyses of current tracing programmes is crucial to inform on the role of contact tracing in protecting the health of communities.

Contributors
LCI conceived of the study. AR and LCI designed the study. ADH, JJ, AE, CXH, and AR were responsible for the search protocol, data extraction, and risk of bias ratings. ADH and JJ wrote the first draft of the manuscript. AE, CXH, AR, and LCI edited the manuscript. All authors had full access to all study data, approved the final version of the manuscript, and agreed to submit for publication. ADH and JJ accessed and verified all data.

Declaration of interests
We declare no competing interests.

Data sharing
All data are available in this Article and in the appendix.

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