Internet interventions for adult illicit substance users: a meta-analysis

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ABSTRACT

Background and Aims  Research has shown that internet interventions can be effective for dependent users of various substances. However, less is known about the effects of these interventions on users of opioids, cocaine and amphetamines than for other substances. We aimed to investigate the effectiveness of internet interventions in decreasing the usage of these types of substances.

Methods  We conducted a systematic literature search in the databases of PubMed, PsycINFO, Embase and the Cochrane Library to identify randomized controlled trials examining the effectiveness of internet interventions compared with control conditions in reducing the use of opioids, cocaine and amphetamines. No setting restrictions were applied. The risk of bias of the included studies was examined according to the Cochrane Risk of Bias assessment tool. The primary outcome was substance use reduction assessed through toxicology screening, self-report or both at post-treatment and at the follow-up assessment.

Results  Seventeen studies with 2836 adult illicit substance users were included. The risk of bias varied across the included studies. Internet interventions decreased significantly opioid [four studies, $n=606$, $g=0.36$; 95% confidence interval (CI) = $0.20$–$0.53$, $P < 0.001$] and any illicit substance use (nine studies, $n=1749$, $g=0.35$; 95% CI = $0.24$–$0.45$, $P < 0.001$) at post-treatment. Conversely, the effect of internet intervention for stimulant users was small and non-significant (four studies, $n=481$, $P=0.164$). Overall, internet interventions decreased substance significantly use at post-treatment (17 studies, $n=2836$, $g=0.31$; 95% CI = $0.23$–$0.39$, $P < 0.001$) and at the follow-up assessments (nine studies, $n=1906$, $g=0.22$; 95% CI = $0.07$–$0.37$, $P=0.003$).

Conclusions  Internet interventions demonstrate small but significant effects in decreasing substance use among various target populations at post-treatment and at the follow-up assessment. However, given the small number of available studies for certain substances, the findings should be interpreted with caution.

Keywords  Internet interventions, meta-analysis, opioid, stimulant, substance use, web-based.

INTRODUCTION

Illicit substance use is a major health issue associated with serious physical, psychological and social harm [1]. Approximately 187 000 of 27 million illicit substance users die annually due to drug-related deaths [2]. Causes of death include drug overdoses and delayed or chronic drug-related medical consequences, including infections, fatal liver diseases, respiratory and cardiovascular diseases. The most widely used illicit substances are cannabis (although recently decriminalized/legalized in various countries) opioids, amphetamines and cocaine, of which the global use estimates are 3.9, 0.7, 0.7 and 0.4%, respectively [2]. Moreover, previous findings indicate that regardless of the primary substance, illicit substance abuse and dependence, referred to currently as substance use disorder, accounts for approximately 20 million disability-adjusted life years (DALYs) lost every year worldwide [3]. This is a substantial amount given that, for example, other serious disorders, such as dementia, account for approximately only 11 million DALYs [4]. Consequently, illicit substance use imposes a significant societal and economic burden [5–7].

Pharmacological interventions are not yet available for amphetamine and cocaine users [8]. Moreover, the literature indicates mixed results for opioid users [9].
However, several meta-analyses have shown that psychosocial and behavioural treatments are effective in reducing illicit substance use [10–15]. Unfortunately, only 20% of individuals with a substance use disorder utilize mental health and addiction services [16]. This phenomenon can be attributed to a variety of reasons, such as lack of availability of treatment services, available but overcrowded programmes, time conflicts, financial barriers, fear of stigma and the requirement of abstinence as a goal [17–19]. Internet interventions might be a novel approach to overcoming these obstacles [20] by increasing the number of people receiving standardized evidence-based treatments, minimizing therapists’ time and decreasing treatment costs [21].

Two distinct internet-based approaches are applied commonly in the substance use literature, unguided stand-alone internet interventions and internet interventions as an add-on to treatment as usual (TAU). These approaches offer distinct advantages: unguided stand-alone internet interventions are capable of supporting numerous substance users simultaneously with a low threshold for accessibility [22–25]. Conversely, add-on internet interventions that are combined with face-to-face support are more intensive treatment in which the support of a mental health professional and the high level of convenience and flexibility of internet interventions are combined. Both approaches have demonstrated encouraging effects for nicotine, alcohol and cannabis users in previous meta-analyses [22,23,26].

However, to the best of our knowledge, no previous meta-analysis has examined the effectiveness of internet interventions regarding substance use reduction in users of opioids, cocaine, amphetamines and any illicit substances [27]. The present meta-analysis aimed to examine to what extent internet interventions are effective in reducing the use of opioids, cocaine, amphetamines and any illicit substances in adults compared to controls. By the term ‘any illicit substance users’, we refer to individuals who use at least one illicit substance and are included in transdiagnostic interventions targeting various substances at once.

**METHODS**

Identification of studies

We conducted a systematic literature search up to January 2016 on the following databases: PubMed, PsycINFO, EMBASE and Cochrane Central Register of Controlled Trials. We used various combinations of key and index terms covering the concepts of substances (drug abuse, addiction, drug dependence, polydrug, heroin, cocaine, crack, opioid, benzodiazepine, ecstasy, amphetamine, methamphetamine) and internet interventions (internet, web, online, computer, mobile) (the full search string for PubMed is given in Supporting information, Appendix S1). Furthermore, we applied a filter for randomized controlled trials (RCTs) in these databases. We conducted additional searches by checking references of the included studies [28]. Our initial selection was based on titles and abstracts. Subsequently, full texts of studies possibly meeting inclusion criteria were retrieved and evaluated. No language restrictions were applied. All searches and screenings were performed independently by two of the authors (N.B.) and (E.K.), and disagreements were resolved through discussion. The identified interventions were either web-based or computerized; however, for the sake of clarity, we will refer to the included interventions as ‘internet interventions’.

Eligibility criteria

We included RCTs that compared internet or computerized interventions with active [e.g. TAU, motivational interviewing (MI), brief intervention (BI), psychoeducation] or non-active (e.g. waiting-list, assessment-only) control conditions. The RCTs had to focus upon adult current users of illicit substances, such as cocaine, amphetamines, opioids or any illicit substances. By ‘users of any illicit substances’, we describe individuals who use at least one illicit substance. Transdiagnostic interventions, targeting those users of any illicit substances, screen for the most common substances, including opioids, cocaine, stimulants, cannabis and alcohol. Furthermore, studies had to include a measurement of substance use of the participants’ at post-test, measured through self-report, toxicology screening or both. No distinction was made between use and abuse or dependence. Studies that did not focus on protocolized interventions but on the internet as a communication medium (e.g. by e-mail, chat or video consultation) were excluded (see Fig. 1).

Quality assessment

The validity of the included studies was assessed according to the criteria of the Cochrane risk of bias assessment tool [29]. We tested (1) adequacy of allocation sequence generation, (2) concealment of the allocation to the particular conditions, (3) blinding of the participants and personnel, (4) blinding of the outcome assessors, (5) appropriate handling of incomplete outcome data by applying an intention-to-treat design, (6) selective outcome reporting, and finally (7) other potential threats to validity, such as early cessation of the trial or extreme baseline imbalances. Two authors (N.B. and E.K.) assessed the risk of bias, and disagreements were resolved by discussion.
Data extraction

We extracted a set of outcome variables measuring the same construct; namely, substance use reduction. This practice enabled us to summarize available findings adequately. In case the studies reported more than one relevant outcome, we aggregated the means of these variables to yield an overall mean effect size for each study. These variables included: (a) mean maximum number of days or weeks abstinent throughout treatment, (b) number of positive urine samples, (c) consumption within previous weeks or months and (d) post-treatment and follow-up scores based on self-report scales. Data were extracted for both post-treatment and the longest follow-up (6 months post-randomization and beyond) outcomes available, and these were analysed separately. Additional aspects of the studies were coded, including participant characteristics (primary substance, recruitment method, setting, medication); intervention characteristics (type of intervention, format, duration, type of assessment, inclusion and exclusion criteria, number of sessions, attrition); control characteristics; length of the follow-up period; and the country in which the trial was carried out (see Table 1). In cases where eligible studies did not report the necessary data to conduct quantitative analyses, we attempted to contact the first author to gain the necessary data [30–32]. In the event of no reply, the studies were not included in the present systematic review [30].

Meta-analyses

We chose to calculate Hedges’ $g$ as the effect size indicating the difference for each comparison between an internet intervention and a control condition. The effect size for small sample bias was corrected according to the procedures suggested by Hedges & Olkin [33]. Given that there is no gold standard of how to report results in the substance use literature [34], we expected to encounter a great variety of outcome reporting styles. The computer program Comprehensive Meta-Analysis (CMA, version 3.3.070) was used for all analyses. Effect sizes were calculated by (a) subtracting the average post-treatment score of the intervention group from the mean score of the control group and dividing the result by the pooled standard deviations of the two groups [35], (b) transforming test statistics (e.g. $t$, $F$, $r$) into the standardized mean difference [36], (c) transforming odds ratios into the standardized mean difference [37]; and (d) transforming medians and interquartile range into means and standard deviations [38]. Effect sizes of approximately 0.8 can be considered large, 0.5 as moderate and 0.2 small [35]. When studies included two or more intervention conditions, we split the control condition into two or more groups and divided the

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| Study          | Primary substance | Measure | Recruiting        | Inclusion criteria                                                                 | Internet format | Intervention       | Comparison | Duration | NSession | Medic. | Overall attrition | Average age & SD |
|---------------|-------------------|---------|-------------------|------------------------------------------------------------------------------------|----------------|---------------------|------------|----------|-----------|--------|-------------------|------------------|
| Bickel, 2008  | Opioids           | U       | Community         | M/F, DSM-IV opioid dependence                                                      | ADD            | CRA + CM            | TAU        | 23       | 72        | Bupren. | 40%               | 28.6 ± 9.1       |
| Christensen, 2014 | Opioids           | U       | Community         | M/F, DSM-IV opioid dependence                                                      | ADD            | CRA + CM + TAU      | CM + TAU   | 12       | 36        | Bupren. | 27%               | 34.4 ± 9.9       |
| Chopra, 2009  | Opioids           | U       | Community         | M/F, DSM-IV opioid dependence                                                      | ADD            | CRA + CM + TAU      | TAU        | 12       | 48        | Bupren. | 27%               | 31.8 ± 10.2      |
| Marsch, 2014  | Opioids           | U       | Out-patient clinic| M/F, DSM-IV opioid dependence                                                      | ADD            | CRA + TAU           | TAU        | 52       | 30        | Meth.  | 61%               | 40.7 ± 9.8       |
| Brooks, 2010  | Stimulants        | S,U     | Out-patient clinic| M/F, DSM-IV cocaine abuse/dependence                                               | ADD            | CRA + CM + TAU      | TAU        | 8        | 48        | NR     | 8%                | 43.1 ± 9.4       |
| Carrol, 2014  | Stimulants        | U       | Out-patient clinic| M/F, DSM-IV cocaine dependence                                                     | ADD            | CBT + TAU           | TAU        | 8        | 8         | Meth.  | 26%               | 42 ± 9.6         |
| Schaub, 2012  | Stimulants        | S       | Community         | M/F, cocaine use ≥ 3 times past 30 days                                            | SA             | CBT                 | EDUC.      | 6        | 8         | NO     | 85%               | 34.2 ± 8.8       |
| Tait, 2014    | Stimulants        | S       | Community         | M/F, ATS use in past 90 days                                                       | SA             | CBT                 | WLC        | 12       | 3         | NO     | 50%               | 22.4 ± 6.3       |
| Campbell, 2014| Any               | U       | Out-patient clinic| M/F, illicit substance use past 90 days                                             | ADD            | CRA + CM + TAU      | TAU        | 12       | 48        | NO     | 56%               | 34.9 ± 10.9      |
| Carrol, 2008  | Any               | S,U     | Out-patient clinic| M/F, any DSM-IV substance use past 60 days                                          | ADD            | CBT + TAU           | TAU        | 8        | 6         | NR     | 34%               | 41.6 ± 10.2      |
| Fals-Stewart, 2010 | Any               | S,U     | Residential center| M/F, any DSM-IV substance use past 60 days                                          | ADD            | CR + TAU            | TAU + TT   | 8        | 24        | NR     | 8%                | 32.8 ± 6.9       |
| Christoff, 2015| Any               | S       | Community         | M/F, ASSIST between 4 and 26                                                        | SA             | MI                  | MI         | 1D       | 1         | NR     | NA                | 24 ± 5.4         |
| Schwartz, 2014| Any               | S,H     | Hospital          | M/F, ASSIST between 4 and 26                                                        | SA             | BI                  | BI         | 1D       | 1         | NR     | NA                | 36.1 ± 14.7      |
| Ondersma, 2005| Any               | S       | Hospital          | F, any illicit drug use 30 days prior pregnancy                                     | SA             | MI                  | AO         | 1D       | 1         | NR     | NA                | 23.4 ± 4.9       |
| Ondersma, 2007| Any               | S,U,H   | Hospital          | F, any illicit drug use 30 days prior pregnancy                                     | SA             | MI                  | AO         | 1D       | 1         | NR     | NA                | 25.1 ± 5.6       |
| Ondersma, 2014| Any               | S,U,H   | Hospital          | F, any illicit drug use 30 days prior pregnancy                                     | SA             | MI                  | AO         | 1D       | 1         | NR     | NA                | 26.6 ± 6         |
| Sinadinovic, 2012| Any               | S       | Community         | M/F, DUDIT over zero                                                               | SA             | MI                  | AO         | 1D       | 1         | NR     | 66%               | 32.6 ± NR        |

ADD = add-on intervention; ASSIST = Alcohol, Smoking and Substance Involvement Screening Test; ATS = amphetamine-type stimulants; BI = brief intervention; Bupren. = buprenorphine; CBT = cognitive–behavioural therapy; CM = contingency management; CR= cognitive rehabilitation; CRA = community reinforcement approach; D = day; DUDIT = Drug Use Disorders Identification Test; F = female; M = male; Medic. = medication; Meth. = methadone; MI = motivational interviewing; NR = not reported; EDUC = psychoeducation; SA = standalone intervention; SD = standard deviation; TAU = treatment as usual; TT = typing tutorial; W = week; WLC = waiting list control; NA = not available.

aBy the term 'any illicit substances' we describe individuals that use at least one illicit substance and are included in transdiagnostic interventions targeting various substances at once. bBy the term 'measure' we describe the primary outcome variable. This column also indicates which study provided multiple measures to our analyses; H = hair analyses; S = self-report; U = urine analyses.
sample size by that number [39]. Consequently, intervention groups were compared separately with the relevant control conditions. If the author reported results separately for certain subgroups (e.g. gender or occupational status), the subgroups were combined and compared as one group with the relevant control condition [39,40]. As we expected heterogeneity among the studies, we decided to calculate mean effect sizes using a random-effects model. This model assumes that the included studies were drawn from populations of studies that differed systematically from one another [41]. We calculated the $I^2$ statistic as an indicator of heterogeneity in percentages to test the homogeneity of effect sizes. A value of 0% suggests no observed heterogeneity, while 25, 50 and 75% suggest low, moderate and high heterogeneity, respectively. Furthermore, we estimated 95% confidence intervals (CI) around $I^2$ [42], using the non-central $\chi^2$-based approach within the HETEROGI module for Stata [43].

Publication bias was tested by inspecting the funnel plot visually. Furthermore, Egger’s linear regression test of the intercept was applied to examine if the bias captured by the funnel plot was significant [44]. We also used Duval & Tweedie’s trim-and-fill procedure [45], which produces an imputed estimate of the effect size accounting for missing studies. We investigated the presence of outliers by examining whether effect sizes and 95% CI of each study overlapped with the 95% CI of the pooled effect size: if an outlying effect size was identified, it was excluded to examine the extent to which it affects the results. Factors that, according to the literature, may have led to heterogeneity and differences between the results of individual studies were investigated through subgroup analyses. Various subgroup analyses were conducted (see Table 3, Supporting information, Appendix S2), according to the mixed-effects model [41]. In this model, studies within subgroups are pooled with the random-effects model, whereas analyses for significant differences between subgroups are conducted with the fixed-effects model. Given the similar behavioural and physiological effects of cocaine and amphetamines [46] and the low number of available studies for each substance, we decided to combine the RCTs targeting these populations in the subgroup analyses creating a combined group; namely, stimulant users. Finally, we conducted three univariable meta-regression analyses to assess the association of the effect size of the internet interventions on substance use reduction (as reported at post-test) with (a) the duration, (b) the number of sessions and (c) the risk of bias of the assessed studies. In a meta-regression analysis, the outcome variable is being predicted according to explanatory variables. The resulting regression coefficient describes how the outcome variable changes with a unit increase in the explanatory variable [47].

**Power calculation**

As we expected to find only a limited number of studies for each primary substance, we conducted a power calculation to investigate the necessary amount of studies that would have to be included to have sufficient statistical power to determine relevant post-treatment and follow-up effects [41]. Assuming a small effect size of $g = 0.30$, with a moderate level of between-study variance ($\tau^2$), a statistical power of 0.80 and an alpha of 0.05, we estimated that five studies with a mean of 60 participants per condition would be required, or 10 studies with 30 participants per condition. Alternatively, we would need 10 studies with 17 participants per condition to detect an effect size of $g = 0.4$, or five studies with 34 participants per condition.

**RESULTS**

**Selection and inclusion of studies**

From the 3057 abstracts (2138 after removal of duplicates), we retrieved 580 full-text papers for possible inclusion in the present systematic review, 563 of which were excluded because they did not meet the inclusion criteria. The flowchart describing the inclusion process is presented in Fig. 1. We included in the analysis a total of 17 individual studies with 18 individual comparisons that met all the inclusion criteria.

**Characteristics of included studies**

Selected characteristics of the included studies are presented in Table 1. In the present meta-analysis, 17 studies with a total of 2836 participants were included ($n = 1461$ in the intervention condition; $n = 1375$ in the control condition). Eight studies included participants based on a DSM-IV diagnosis, six studies included individuals according to a cut-off score based on self-reported substance use, two studies included individuals with an elevated substance score on the Alcohol, Smoking and Substance Involvement Screening Test (ASSIST) [48] and one study included participants who scored more than zero at the Drug Use Disorders Identification Test (DUDIT) [49]—indicating at least minimal illicit substance use. Three target populations were addressed in these trials; namely, opioid users ($n = 4$), stimulant users ($n = 4$) and users of any illicit substances ($n = 9$). By ‘users of any illicit substances’, we describe individuals who use at least one illicit substance. RCTs, which target users of any illicit substances, screen for the most common substances, including opioids, cocaine, stimulants, cannabis and alcohol, and provide transdiagnostic interventions. The majority of the studies included recruited the participants from a clinical setting (out-patient facility $n = 5$, hospital $n = 4$, residential centre $n = 1$), whereas the remaining
seven studies recruited their participants from the community. Gender was distributed equally between the intervention and control conditions. The overall study attrition varied substantially, ranging from 8 to 89%. Five of the included studies provided medication to the patients [32,39,50–52], three specifically excluded patients receiving medication [31,53,54], and the remaining studies did not report the proportion of patients on medication [40,55–62]. Nine trials applied an add-on intervention, whereas eight trials applied an unguided standalone intervention. Three of the nine studies that included users of any illicit substances compared unguided standalone internet interventions with non-active controls, three compared add-on interventions with TAU only and three compared unguided standalone internet interventions with other types of active controls (MI, BI, psychoeducation). The four opioid primary substance studies compared add-on interventions with TAU only. Of the four studies on stimulant users, two compared add-on interventions with TAU only and two compared an unguided standalone intervention with a non-active control.

The interventions applied varied according to the target population (interventions were defined as described in Table 2). The community reinforcement approach (CRA) was applied for opioid users, whereas cognitive behaviour therapy (CBT) was the prevailing approach for stimulant users. Internet interventions targeting users of any illicit substances varied considerably in terms of treatment approach and number of sessions, with MI being the dominant approach. Of the 17 studies in total, some studies employed more than one type of assessment. Specifically, 12 employed measures based on toxicology screenings, such as urine or hair analyses, and 11 studies employed self-report measures, such as the ASSIST, the DUDIT and through online questionnaires that measured number of days using. The studies were carried out in five different countries (Australia n = 1, Brazil n = 1, Sweden n = 1, Switzerland n = 1, and the United States n = 13).

Quality assessment

The methodological quality of the studies varied (see Supporting information, Appendix S3). Fifteen of the 17 studies reported an adequate sequence generation. Six studies described adequate allocation concealment. None of the studies blinded the participants and personnel and were hence considered as at high risk of bias. Thirteen of the studies employed toxicology screenings to measure substance use, thus blinding of the outcome assessment was considered as at low risk of bias, while four studies employed self-reporting scales, thus blinding of the outcome assessment was considered as at high risk of bias for these studies. Incomplete outcome data were handled correctly in 13 studies. All the studies reported all expected outcomes, and finally, none of the included studies had other potential threats to validity. Three studies met six of the seven quality criteria, 10 met five criteria, two met four criteria and the remaining two studies met only three criteria. However, only three of the included RCTs provided their protocol online for evaluation [31,53,54]. As a result, it is impossible to investigate protocol violations, which have been shown to be a serious threat to the validity of the study results [63].

The overall effect of internet interventions on illicit substance users

Because the number of studies per primary substance was small, we opted to pool the studies together to increase the statistical power and consequently enable the investigation of study characteristics in subgroup analyses. There was a small but significant overall effect of internet interventions on substance use reduction when all experimental conditions (n = 18) were compared with the control conditions at post-treatment assessment. Heterogeneity was moderate. A visual inspection of the forest plot indicated two possible outliers [57,58], in which the effect size did not overlap with the 95% CI of the pooled effect size. Therefore, we excluded the possible outliers. This exclusion resulted in a marginal increase of the effect size in favour of internet interventions (see Fig. 2). A post-hoc power calculation indicated that our set of studies had sufficient statistical power to detect an effect size of 0.16 on the basis of the random-effects model (with a low level of between-study variance, $r^2 = 0$, a statistical power of 0.80 and a significance level of $P < 0.05$). Inspection of the funnel plot and Duval & Tweedie’s trim-and-fill procedure did not indicate publication bias. Moreover, nine studies reported follow-up outcomes of internet interventions against control conditions, six of which were studies including users of any illicit substances and three included stimulant users. At the follow-up (6–12 months), internet interventions (n = 9) outperformed control groups on substance use reduction for illicit substance users. The effect size was small but significant (n = 1906, g = 0.22; 95% CI = 0.07–0.37; $P = 0.003$). A post-hoc power calculation indicated that our set of studies for follow-up effects had sufficient statistical power to detect an effect size of 0.22 on the basis of the random-effects model (with a low level of between-study variance, $r^2 = 0.001$, a statistical power of 0.80 and a significance level of $P < 0.05$) (Fig. 2 and Table 3).

Substance-specific results of internet interventions

The effects of internet interventions on substance use reduction were similar for opioid users (n = 606, g = 0.36; 95% CI = 0.20–0.53, $P < 0.001$) and users of
any illicit substances, indicating a small but significant effect. A visual inspection of the forest plot indicated two possible outliers [57,58], in which the effect size did not overlap with the 95% CI of the pooled effect size. Therefore, we excluded the possible outliers. This exclusion resulted in a marginal increase of the effect size for users of any illicit substances in favour of internet interventions (n = 1749, g = 0.35; 95% CI = 0.24–0.45, P < 0.001; I² = 0; 95% CI = 0–71, P = .750). Conversely, the effect of internet intervention on stimulant use reduction was small and non-significant when compared to the control conditions (n = 481, P = 0.164) (Fig. 2 and Table 3).
Subgroup analyses

We conducted a series of subgroup analyses (see Table 3, Supporting information, Appendix S2). These analyses indicated that add-on interventions were more effective compared to unguided standalone interventions and the type of assessment; specifically, toxicology screenings were associated with higher effect sizes compared to self-report. Interventions delivered in an out-patient clinic via computer were significantly more effective compared to interventions delivered via computer at a university and via the internet at home. Finally, the type of eligibility screening also affected the outcomes; specifically, DSM-IV diagnoses were associated with higher effect sizes compared to cut-off scores on self-reported substance use (Table 3).

Meta-regression analyses

The meta-regression analyses did not reveal significant associations between the effect size of internet interventions on substance use reduction and (a) duration (slope: 0.002; 95% CI = −0.007 to 0.012, \( P = 0.652 \)), (b) number of sessions (slope: 0.003; 95% CI = −0.003 to 0.008, \( P = 0.338 \)) and (c) risk of bias (slope: −0.063; 95% CI = −0.187 to 0.062, \( P = 0.326 \)).

DISCUSSION

The present systematic review was the first, to our knowledge, to examine the effectiveness of internet interventions regarding substance use reduction of opioids, stimulants and any illicit substances at both post-treatment and at the follow-up. We found a small but significant overall effect size of internet interventions for the reduction of illicit substance use at post-treatment (\( g = 0.30 \)) and at the follow-up assessment (\( g = 0.22 \)). Specifically, internet interventions decreased opioid use (\( g = 0.36 \)) and any illicit substance use significantly (\( g = 0.32 \)), but did not reduce stimulant use significantly. These findings are in line with the broader literature of internet interventions on nicotine, alcohol and cannabis use reduction that indicate small but significant effects [22,23,26].

Our subgroup analyses displayed varying effects, depending on certain characteristics of the trials; namely, format of the intervention, type of the assessment, screening and setting. Specifically, (a) add-on interventions displayed higher effects compared to unguided standalone interventions, (b) studies which applied toxicity screenings as outcome measure were associated with higher effect sizes compared to studies that applied self-report measures, (c) DSM-IV diagnoses were associated with higher effect sizes compared to cut-off scores on self-reported substance use, and finally (d) studies that conducted the intervention in clinical out-patient clinics showed higher effect sizes compared to other settings.

Limitations

We were not able to extract data from all eligible studies, because one study did not report outcome data [27]. This study found no significant effects in substance use when adding an add-on internet intervention to the TAU setting. Moreover, it should be taken into account that certain characteristics of the trials (add-on intervention, toxicity...
screening, out-patient clinic) are associated predominantly with intensive treatments that were compared with TAU, while other characteristics (e.g. unguided standalone interventions, self-report assessments, home setting) were associated predominantly with less intensive interventions and compared primarily (except for three [40,58,63] studies) with inactive interventions, such as assessment-only and waiting-list. For this reason, we suggest that the unexpected finding of toxicology screenings and DSM diagnoses being associated with higher effect sizes might be the fact that they are applied in studies with intensive treatments that target severe users and have higher chances of yielding high effect sizes compared to studies with less intensive treatments that target less severe users. Another explanation might be that a study requiring toxicology screenings might lead to selection bias by recruiting more adherent patients. Therefore, these findings should be considered with caution. Moreover, five of the included studies applied completers’ only analyses [31,39,50,56,58], which might have induced heterogeneity in our findings, especially considering the high attrition of certain studies. Also, we experienced some difficulty in assessing the influence medication might have had on the effect sizes, as only a minority of studies who use at least one illicit substance and are included in transdiagnostic interventions targeting various substances at once. Active controls include: TAU, MI, BI, EDUC; inactive controls include: WLC, add-on.

| Subgroup analyses of associations between effect sizes and study characteristics (Hedges’s g). |
|----------------------------------|-----------------|-----------------|-----------------|-----------------|-----------------|
|                                    | Ncomp | g      | 95% CI | I² | 95% CI | Pb       |
| All studies                        | 18    | 0.30   | 0.19 to 0.41*** | 50 | 13–71 |
| 2 possible outliers removed         | 16    | 0.31   | 0.23 to 0.39*** | 0  | 0–52  |
| Primary substance                  |       |        |        |     |       |
| Anyd                              | 9     | 0.32   | 0.15 to 0.49*** | 69 | 38–85 | 0.146   |
| Opioids                           | 5     | 0.36   | 0.20 to 0.53*** | 0  | 0–85  |
| Stimulants                        | 4     | 0.13   | -0.05 to 0.31 | 0  | 0–85  |
| Subgroup analyses (n = 18)         |       |        |        |     |       |
| Control group                      |       |        |        |     |       |
| Active                            | 13    | 0.31   | 0.16 to 0.46*** | 62 | 31–79 | 0.978   |
| Non-active                        | 5     | 0.31   | 0.17 to 0.45*** | 0  | 0–79  |
| Type                              |       |        |        |     |       |
| CRA + CM                          | 6     | 0.39   | 0.26 to 0.52*** | 0  | 0–75  | 0.382   |
| MI                                | 5     | 0.30   | 0.16 to 0.44*** | 0  | 0–79  |
| CBT                               | 4     | 0.19   | 0.02 to 0.35  | 17 | 0–87  |
| Other                             | 3     | 0.34   | -0.18 to 0.85 | 90 | 74–96 |
| Format                            |       |        |        |     |       |
| Add-on                            | 10    | 0.41   | 0.30 to 0.52*** | 12 | 0–54  | 0.011   |
| Standalone                        | 8     | 0.17   | 0.03 to 0.32  | 43 | 0–75  |
| Type of assessment                |       |        |        |     |       |
| Tox. screening                    | 12    | 0.42   | 0.32 to 0.52*** | 0  | 0–58  | 0.016   |
| Self-report                       | 11    | 0.26   | 0.05 to 0.42*  | 71 | 46–84 |
| Screening                         |       |        |        |     |       |
| DSM-IV                            | 9     | 0.42   | 0.27 to 0.56*** | 22 | 0–63  | 0.048   |
| Cut-off scores                    | 9     | 0.21   | 0.07 to 0.36*  | 55 | 4–79  |
| Medication                        |       |        |        |     |       |
| Yes                               | 6     | 0.34   | 0.19 to 0.48*** | 0  | 0–75  | 0.774   |
| No                                | 3     | 0.22   | -0.06 to 0.50 | 67 | 0–91  |
| NR                                | 9     | 0.30   | 0.11 to 0.50**  | 67 | 34–84 |
| Analyses                          |       |        |        |     |       |
| ITT analyses                      | 13    | 0.31   | 0.20 to 0.42*** | 36 | 0–67  | 0.883   |
| Comp. analyses                    | 5     | 0.33   | 0.04 to 0.63*  | 69 | 20–88 |
| Recruitment                       |       |        |        |     |       |
| Clinical                          | 10    | 0.34   | 0.18 to 0.49*** | 63 | 28–82 | 0.341   |
| Community                         | 8     | 0.23   | 0.01–0.37**    | 11 | 0–71  |
| Setting                           |       |        |        |     |       |
| Computer, out-patient clinic      | 9     | 0.36   | 0.26 to 0.47*** | 0  | 0–65  | 0.002   |
| Computer, hospital                | 4     | 0.23   | -0.44 to 0.50  | 71 | 19–90 |
| Internet, home                    | 3     | 0.11   | -0.07 to 0.29 | 0  | 0–90  |
| Computer, university              | 1     | 0.12   | -0.33 to 0.56 | 0  | NA    |
| Computer, residential centre      | 1     | 0.82   | 0.50 to 1.14*** | 0  | NA    |
| Female-only studies               |       |        |        |     |       |
| Yes                               | 3     | 0.37   | 0.20 to 0.55*** | 0  | 0–90  | 0.448   |
| No                                | 15    | 0.29   | 0.16 to 0.42*** | 57 | 23–76 |

**BI** = brief intervention; **CET** = cognitive enhancement therapy; **Comp.** = completers; **EDUC** = psychoeducation; **ITT** = intention-to-treat; **MI** = motivational interviewing; **NA** = not applicable; **Ncomp** = number of comparisons; **NR** = not reported; **PA** = positive affect; **TAU** = treatment as usual; **Tox.** = toxicology; **WLC** = waiting-list control. "According to the random-effects model; "the P-values in this column indicate if the difference between the effect sizes in the subgroups is significant; CI = confidence interval (*P* ≤ 0.05; **P* < 0.01; ***P* ≤ 0.001); "[57,58]. By the term ‘any illicit substances’ we describe individuals who use at least one illicit substance and are included in transdiagnostic interventions targeting various substances at once. Active controls include: TAU, MI, BI, EDUC; inactive controls include: WLC, add-on."
interventions. Previous research has shown that transdiagnostic interventions for substance users with comorbid mental disorders are more effective [63] and have higher adherence rates [64] compared to separate treatment plans for individual disorders.

Summarizing the empirical literature on the effectiveness of internet interventions for substance users presented several challenges. The utilized outcomes to determine superiority of a condition varied substantially, and although this variation of outcomes has been noted by previous studies [11,12,64] little effort has been taken place to overcome this challenge. Therefore, we suggest that future studies should approach this issue to further improve the comparability of outcome findings. Furthermore, the relatively high rate of attrition might be a serious restriction. This is especially true for unguided standalone interventions and is often considered a serious threat to the integrity of the study results [65]. Therefore, further research should investigate possible predictors to prevent attrition.

**CONCLUSION**

Although more research is necessary to assess the effectiveness of internet interventions in illicit substance users, this method of treatment delivery seems to be a promising solution for achieving substance use reduction in these target populations.

**Declaration of interests**

None.

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Supporting Information
Additional Supporting Information may be found online in the supporting information tab for this article.

Appendix S1 Pubmed search string.
Appendix S2 Definition of subgroup analyses.
Appendix S3 Risk of bias assessment.

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