Resting state functional connectivity provides mechanistic predictions of future changes in sedentary behavior

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Sedentary behaviors are increasing at the cost of millions of dollars spent in health care and productivity losses due to physical inactivity-related deaths worldwide. Understanding the mechanistic predictors of sedentary behaviors will improve future intervention development and precision medicine approaches. It has been posited that humans have an innate attraction towards effort minimization and that inhibitory control is required to overcome this prepotent disposition. Consequently, we hypothesized that individual differences in the functional connectivity of brain regions implicated in inhibitory control and physical effort decision making at the beginning of an exercise intervention in older adults would predict the change in time spent sedentary over the course of that intervention. In 143 healthy, low-active older adults participating in a 6-month aerobic exercise intervention (with three conditions: walking, dance, stretching), we aimed to use baseline neuroimaging (resting state functional connectivity of two a priori defined seed regions), and baseline accelerometer measures of time spent sedentary to predict future pre-post changes in objectively measured time spent sedentary in daily life over the 6-month intervention. Our results demonstrated that functional connectivity between (1) the anterior cingulate cortex and the supplementary motor area and (2) the right anterior insula and the left temporoparietal/temporoccipital junction, predicted changes in time spent sedentary in the walking group. Functional connectivity of these brain regions did not predict changes in time spent sedentary in the dance nor stretch and tone conditions, but baseline time spent sedentary was predictive in these conditions. Our results add important knowledge toward understanding mechanistic associations underlying complex out-of-session sedentary behaviors within a walking intervention setting in older adults.

In 2007 it was estimated that ~ 5.3 million global deaths from non-communicable diseases could have been prevented if people engaged in sufficient levels of moderate-to-vigorous physical activity instead of being insufficiently active1. Compounding this further, global statistics show the prevalence of physical inactivity is increasing2,3. Over a third of the US population (34.8%) lead sedentary lifestyles2–4 and the economic burden caused by physical inactivity is estimated to cost private and public health-care systems $53.8 billion per year5,6.

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To combat the negative consequences of sedentary behaviors, particularly in older adults, the field has studied extensively the beneficial effects of exercise interventions\(^5\)-\(^9\). The most well studied exercise interventions are walking interventions, which are both economical and easily accessible, particularly for older adults\(^10\). These studies have led to numerous discoveries on the beneficial effects of increased walking on cognitive function, particularly, processing speed, memory and executive function\(^11\). Walking interventions also have been shown to increase hippocampal volume\(^12\) and the plasticity of functional brain networks\(^13\). These results are particularly important given that these same outcomes are also associated with age-related decline\(^14-17\). However, engaging in a significant behavioral change is non-trivial and despite significant efforts to understand determinants of sedentary lifestyles, the prevalence of physical inactivity continues to increase\(^2,3\).

Sedentary behaviors are not simply the inverse of moderate-to-vigorous physical activity\(^18,19\). For example, a person can both perform 30 min of moderate-to-vigorous physical activity achieving recommended levels\(^20\) and also engage in a high volume of sedentary behavior throughout the rest of the day. Further, the determinants of sedentary behaviors are distinct from those of physical activity engagement too\(^18\). For example, to engage in a bout of physical activity one must inhibit a desire to minimize effort one time, whereas avoiding sedentary behaviors throughout the day requires consistent awareness and self-regulation of such behaviors\(^21\). Understanding the determinants of sedentary behaviors has relied upon psychological frameworks and cognitive-behavioral theories\(^22\). Automatic processes, attitudes and habits and self-regulation have been suggested to regulate sedentary behaviors\(^28,31\). For example, in rodents, local field potentials in and coherence between the aMCC and the dAI correlate with relative performance on a physical effort-based task\(^36\). In humans, neuroimaging studies have demonstrated that the aMCC is a critical region for decision-making of choices involving motor-costs\(^28\) and further, that activity in the aMCC and the dAI represent the devaluation of rewards associated with physical effort\(^1\). Additionally, these same regions are consistently implicated in inhibitory control\(^17,30\), a higher order executive function shown to be needed to overcome physical effort minimization\(^28\). Together, this theoretical and experimental evidence may suggest a role for the aMCC and the dAI in the regulation of sedentary behaviors.

The discovery of neural predictors of future sedentary behaviors may provide both strong predictive strength as well as mechanistic information relevant for intervention development. The utility and efficacy of functional connectivity (FC) to predict future behavioral outcomes has been demonstrated in previous research. For example, Saghayi and colleagues predicted treatment response in social anxiety disorder with FC, better than clinical measures\(^24\) and colleagues predicted adherence to mental training programs using FC\(^39\) and Whitfield-Gabrieli and colleagues predicted treatment response in social anxiety disorder with FC\(^40\), better than clinical measures alone\(^40\).

The aim of this present study therefore was to evaluate if the FC of two a-priori defined brain regions (aMCC and the r-dAI) implicated in inhibitory control and physical effort decision making, at baseline, could predict future change in objectively measured sedentary behavior in older adults participating in a 6-month randomized controlled trial of exercise (which included a walking, a dancing and a stretching control condition).

**Methods**

**Participants and study design.** This study presents results of a secondary analysis of baseline data from participants who participated in a 6-month randomized controlled exercise trial (clinical study identifier: NCT01472744, November 16, 2011). The study procedures were approved by the University of Illinois Institutional Review Board and written informed consent was obtained from all participants prior to any research activities. All methods were carried out in accordance with the Declaration of Helsinki. Healthy but low active older adults were recruited in Champaign County. Two hundred and forty-seven (169 women) low-active (less than two bouts of self-reported moderate exercise per week within the past 6 months) older adults met inclusion criteria for the initial clinical trial. Of which one hundred and sixty-five underwent functional magnetic resonance imaging (fMRI). Participants in the initial trail were randomized to one of four intervention groups; a walking intervention, a walking intervention plus a dietary supplement, a dancing intervention and a control stretch and toning intervention. For the purpose of this analysis, we combined the two walking groups to increase the sample size as the walking portion of the intervention as identical and no significant differences in outcome measures or demographics was found between these two groups (supplementary material 1). All groups met for approximately one hour three times per week for six months. For this analysis, we excluded participants who...
did not adhere to more than 50% of the intervention sessions (n = 9), for having incomplete accelerometer data available (n = 7), high motion artefact in the IMRI scan (see below for criteria, n = 2), or for influential outlier data points in the outcome variable (see criteria below, n = 4). 143 participants were ultimately included in this study. For more details on this clinical trial, its primary outcomes and neuroimaging data, please refer to earlier work41–44. Initially, to enroll in the study, participants must have met the following criteria: were between the ages of 60 and 80 years old, free from psychiatric and neurological illness and had no history of stroke, transient ischemic attack, or head trauma, scored < 23 on the Mini-Mental State Exam, < 21 on a Telephone Interview of Cognitive Status questionnaire and < 10 on the Geriatric Depression Scale, at least 75% right-handed based on the Edinburgh Handedness Questionnaire (a criterion related to functional magnetic resonance imaging (MRI) analyses), demonstrated normal or corrected-to-normal vision of at least 20/40 and no color blindness, screened for safe participation in an MRI environment (e.g., no metallic implants that could interfere with the magnetic field or cause injury and no claustrophobia) and reported to have participated in no more than two bouts of moderate exercise per week within the past 6 months (with the goal of recruiting low active older adults). Our current analysis asks a novel question of this dataset that has not been previously assessed. Table 1 contains complete characterization of the study participants broken down by each intervention group.

### Table 1. Participant characteristics. Baseline sedentary time = estimated baseline average daily minutes spent sedentary. Post sedentary time = estimated post-intervention average daily minutes spent sedentary. P-value represents the results of ANOVA (continuous) or chi-square test of independence (categorical) tests on outcome and demographic variables between groups.

|                          | Walk          | Dance         | Stretch and Tone | P   |
|--------------------------|---------------|---------------|------------------|-----|
| N                        | 63            | 40            | 40               |     |
| Age (mean (SD))          | 65.33 (4.53)  | 66.15 (4.74)  | 65.72 (4.89)     | 0.683 |
| Baseline sedentary time (mean (SD)) | 537.43 (91.67) | 530.11 (92.84) | 564.79 (75.35) | 0.170 |
| Post sedentary time (mean (SD)) | 555.72 (107.9) | 547.24 (83.03) | 574.90 (72.30) | 0.388 |
| Female sex (%)           | 46 (71.9)     | 28 (70.0)     | 28 (70.0)        | 0.970 |
| Increase in sedentary time (%) | 38 (60.3)    | 24 (60.0)     | 22 (55.0)        | 0.851 |

Accelerometry. Time spent sedentary was measured using an ActiGraph accelerometer device (Model GT1M or GT3X; ActiGraph, Pensacola, FL) for one week at baseline and one-week post-intervention. Participants were instructed to wear the accelerometer on the nondominant hip during waking hours for seven consecutive days. For data reduction, the following criteria were applied to the raw data recorded by each monitor: wear time validation criterion of ≥ 10 h of wear time per day for at least 3 days and an interruption period of 60 min. These data were downloaded as activity counts, which represent raw accelerations summed over a specific epoch length (e.g., 1 s) and subsequently processed into activity intensities in ActiLife software package (Version 6; Actigraph, Pensacola, FL). A low intensity proxy for sedentary behavior was derived using older adult-specific cut points such that 50 or fewer counts per minute corresponded with sedentary behavior. Estimated average daily minutes spent in the sedentary category (< 50 counts/min) were calculated by dividing the number of minutes spent in that category by the total number of valid days worn per participant. Our outcome measure (change in time spent sedentary) was calculated as post-test minus pre-test of the estimated average daily minutes spent sedentary.

Magnetic resonance imaging: preprocessing. Participants underwent an MRI scanning session in a 3 Tesla Siemens TIM Trio system with a 12-channel head coil. High-resolution structural MRI scans were acquired using 3D MPRAGE T1-weighted sequences (TR = 1900 ms; TE = 2.32 ms; TI = 900 ms; flip angle = 9°; matrix = 256 × 256; FOV = 230 mm; 192 slices; resolution = 0.9 × 0.9 × 0.9 mm; GRAPPA acceleration factor 2). One run of T2*-weighted resting state echo-planar imaging (EPI) data was obtained with the following parameters: (6 min, TR = 2 s, TE = 25 ms, flipangle = 80°, 3.4 × 3.4 mm² in-plane resolution, 35 4 mm-thick slices acquired in ascending order, Grappa acceleration factor = 2, 64 × 64 matrix).

Preprocessing of the functional resting state data was performed using the CONN-toolbox v.19c, relying upon SPM v.12 (Wellcome Department of Imaging Neuroscience, UCL, London, UK) in MATLAB R2019a (The MathWorks Inc, Natick, MA, USA). The latest default preprocessing pipeline implemented in Conn was performed which consists of the following steps: functional realignment and unwarping, slice timing correction, outlier identification, segmentation (into grey matter, white matter and cerebrospinal fluid) and normalization into standard Montreal Neurologic Institute (MNI) space resampled to 2 mm isotropic voxels for functional data and 1 mm for anatomical data, using 4th order spline interpolation. Functional scans were spatially smoothed using a 6 mm FWHM Gaussian kernel. During the outlier detection step, acquisitions with framewise displacement above 0.9 mm or global BOLD signal changes above 5 standard deviations were flagged as outliers using the Artefact Detection Tools (www.nitrc.org/projects/artifact_detect). Two participants were removed from the final analyses for having > 30 scan volumes flagged. This cut off was determined based on preserving at least 5 min of scanning time. Additionally, mean motion (framewise displacement) was used as a covariate of no interest in all second level analyses. This was done to be over conservative given previous studies have shown high degree of motion-behavior correlations, despite the fact that no motion parameter was significantly correlated with sedentary time in our data (P > 0.05). Denoising of the functional data was performed using a principal
component analysis-based correction method, CompCor. Linear regression was used to remove the effects of these artifacts on the BOLD time series for each voxel and each subject taking into account noise components from cerebral white matter and cerebrospinal fluid, estimated subject-motion parameters (3 rotation and 3 translation parameters and 6 other parameters representing their first order time derivatives), scrubbing (one noise component for each outlier scan detected in the outlier detection step) and constant and first-order linear session effects. Temporal band-pass filtering (0.008–0.09 Hz) was applied to remove physiological, subject-motion and outlier-related artefacts. MRI quality control measures are found in the supplementary material.

Seed-based correlations. The average time series in two regions of interest (ROI), the anterior mid-cingulate (aMCC) and the right dorsal anterior insula (r-dAI) were extracted. We defined our seeds using the 100-parcel functional atlas by Schaefer 2018. Because the functional parcels of the aMCC and the r-dAI extend outside of the anatomical boundaries of interest, we limited our seed ROIs to just the functional parcel constrained by the anatomical boundaries of the aMCC and the r-dAI set by the Harvard–Oxford anatomical atlas. This was done by binarizing the parcels from each atlas and using 'fslmaths' functions (Functional Magnetic Resonance Imaging of the Brain's Software Library, http://www.fmrib.ox.ac.uk/fsl) to multiply the two parcels together (see Fig. 1 for an illustration of the seed ROIs). Then, Pearson’s correlation coefficients were computed between the average time series in each ROI and the time series of all other voxels in the brain and converted to normally distributed z-scores using Fisher transformation prior to performing the second-level general linear model. Individual change in sedentary time was entered as a covariate of interest in the second-level analysis, controlling for nuisance variables, age, gender, baseline sedentary time and mean framewise displacement, in separate general linear models for each ROI. In a confirmatory step, results in this second level analyses were estimated using a height threshold (voxel level \( P < 0.001 \)) and a family-wise corrected cluster-extent threshold (\( P \text{ FWE} < 0.05 \)) and can be found in the supplementary materials.

Statistical analyses. The effect of each intervention on time spent sedentary was assessed using repeated measures analysis of variance. Differences in outcome and demographic variables between groups were assessed.
using analysis of variance for continuous outcomes and chi-square test of independence for categorical variables within the `Table1` function in R. The Breusch-Pagan Test of Heteroskedasticity was performed to ensure homogeneity of variance.

To assess whether baseline measure of sedentary time predicted change in time spent sedentary we ran independent linear regression models using leave-one-out cross validation (LOOCV) in each group with age and sex, (and baseline sedentary time in the FC models). Model assumptions for linear models were checked using Q-Q and fitted vs. residual plots in R. The significant influence of outliers was checked using Cook's distance with a cut off of 0.5 (n = 3 for the stretch and tone group and n = 1 for the walking group).

To test whether seed-based functional connectivity predicted change in time spent sedentary we implemented a nested cross validation procedure. Each outer-layer LOO iteration used data from N-1 subjects to: (a) first select the largest cluster of voxels showing significant ($P < 0.001$) voxel-level associations with time spent sedentary; (b) run an inner-layer cross-validation procedure to fit a linear model between average connectivity in that cluster and time spent sedentary; and (c) compute the average connectivity within this cluster for the left-out subject and use the estimated linear model parameters to predict time spent sedentary for this same left-out subject.

Model performance is presented as cross-validated R$^2$ values. We also present the average prediction error (RMSE) which represents the difference between the observed and predicted values. Statistical significance of the prediction models was assessed via 1000 nonparametric permutations and the p-value of the permutation tests were calculated as the proportion of sampled permutations that are greater or equal to the true prediction correlation.

LOOCV of the seed-based correlation clusters was performed in MATLAB using the "spm_nestedcrossvalidation" code and all other statistics performed in RStudio Version 3.6.3 (R Foundation for Statistical Computing, Vienna, Austria) using "tidyverse"51, "Caret"52 and base R packages.

**Ethics approval.** The University of Illinois Institutional Review Board approved all procedures used in the study.

**Consent to participate.** All participants gave written informed consent before participation in any study procedures, all of which conformed to the Declaration of Helsinki for research involving human subjects.

**Consent for publication and author responsibilities.** All authors agree to the contents of this manuscript and give consent for its publication.

**Results**

One-hundred and forty-three low-active healthy older adults were included in this study. Table 1 outlines participant demographics broken down by intervention condition. The distribution of the change in time spent sedentary (Fig. 2) revealed that a higher proportion of participants increased their time spent sedentary over the course of the intervention with no significant differences in this proportion between intervention conditions (Table 1). No main effect of condition assignment ($F(1, 198) = 1.981, P = 0.167$), time ($F(1, 1) = 2.934, P = 0.087$) or time by condition interaction ($F(1, 13) = 0.137, P = 0.711$) was found for time spent sedentary over the course of the intervention.

**Baseline time spent sedentary.** Baseline time spent sedentary predicted change in time spent sedentary in the stretch and tone and dance groups, but not the walking group (Table 2).

**Functional connectivity.** In the dance and stretch and tone groups, baseline functional connectivity of the aMCC and the r-dAI was not predictive of change in time spent sedentary. In the walking group baseline functional connectivity between the aMCC and the M1/SMA predicted change in time spent sedentary (Table 2 and Fig. 3). Similarly, baseline functional connectivity between the r-dAI and the left temporoparietal/temporoccipital region (areas spanning the middle temporal gyrus, angular gyrus and lateral occipital cortex predicted change in time spent sedentary (Table 2 and Fig. 3). All results from these second level seed-based correlations were confirmed to hold in a whole-sample association analysis using conventional height-level statistical threshold of $P < 0.001$ and cluster threshold of $P < 0.05$ family wise error corrected (supplementary material 3).

**Discussion**

The aims of the current study were to assess whether baseline functional connectivity of brain regions implicated in executive control and effort-based decision making could provide mechanistic predictions of change in time spent sedentary in older adults participating in a randomized control trial of exercise. In the walking group, participating in the most commonly found exercise intervention in the literature, we found that baseline behavioral measures were not predictive of change in time spent sedentary but functional connectivity of the aMCC and r-dAI were predictive. In the aerobic dance group and the control stretch and tone group, FC was not predictive of change in time spent sedentary, but baseline time spent sedentary was.

While our analysis of the objective measures of time spent sedentary did not reveal any differences between intervention conditions, previous research in this same sample15 demonstrated differences in out-of-session aerobic activity between intervention conditions, suggesting that the determinants of exercise and sedentary behaviors (in older adults participating in an exercise intervention) could differ between intervention types. Our result that FC predicted change in time spent sedentary in the walking group only is potentially in line with the idea that specific interventions may result in contextually different behaviors. Notwithstanding, many aerobic exercise interventions in older adults consist only of an active and a control intervention, making this
Figure 2. Histograms of participant changes in sedentary time over the 6-month interventions. A numerically similar proportion of individuals increased as decreased their time spent sedentary. Gold vertical line represents the mean change, “0” on the x-axis represents no change.

Table 2. Prediction of change in time spent sedentary. All models are performed using leave-one-out cross validation. RMSE root mean square error and represents the differences between the observed and predicted outcomes (the lower the value the better the prediction). All significant models survive multiple comparisons using false discovery rate (supplementary material 4). Statistical significance of the prediction models was assessed via 1000 nonparametric permutations and the p-value of the permutation tests were calculated as the proportion of sampled permutations that are greater or equal to the true prediction correlation.
conclusion hard to generalize. One possible interpretation of the differential predictiveness between our experimental conditions is perhaps related to statistical power. When we ran a down sampled analysis of the models using a randomly sampled N of 45 in the walking condition, the aMCC-to-M1/SMA relationship found in the wider walking sample was still present (supplementary material 4), but the r-dAI result was not. Interestingly, when running an exploratory analysis with all conditions combined (N = 143), the aMCC result disappears and the r-dAI result seen in the entire walking sample is present (supplementary material 5). This is perhaps suggestive of a lack of statistical power in the dance and stretch and tone samples, respectively, to detect a relationship between the FC of the r-dAI seed and change in time spent sedentary. Indeed, our power analysis in these intervention condition groups (supplementary material 6) found that we only have 64% power to detect an effect of the size seen in the walking condition. However, at the same time, these exploratory analyses lead us to speculate that the aMCC-to-M1/SMA FC result is perhaps specific to the walking condition. One possible explanation of the walking condition-specific result could lie in a previous analysis of this sample where participants in the walking group self-reported a reduction in the amount of out-of-session aerobic activity across the course of the 6-month intervention whereas those participating in the dance and stretch and tone conditions maintained their aerobic activity levels. In that prior analysis, perceived intensity of the intervention sessions was associated with out-of-session aerobic activity, whereby higher perceptions of session activity were found for the walking group compared to the dance group and thus it was concluded that those in the walking group may have deemed the 3 times per week sessions as sufficient aerobic activity whereas those in the dance and stretch and tone groups may have deemed their session to be necessary but insufficient, leading them to engage in more aerobic activity outside of the intervention sessions. Consequently, given the aMCC’s role in effort-based decision making, it is plausible that this mechanistic prediction of change in time spent sedentary is specific to those engaging in a walking intervention of a given intensity to be perceived as sufficient weekly aerobic exercise.

The main aim of this study was to ask whether resting state functional connectivity could provide mechanistic predictions of change in time spent sedentary. We chose our seed regions (aMCC and the r-dAI) as they have been consistently implicated in effort-based decision making and the integration of motor costs with reward outcomes. Further, these same regions have been implicated in inhibitory control, which has been shown to be important to overcome the posited innate attraction towards effort minimization. The function of the aMCC and its behavioral role has been highly debated (i.e. does it motivate effortful behaviors or engage in decision-making and deployment of cognitive control?). In an attempt to unify these theories, Holyrood and Yeung (2012) proposed that the aMCC supports the selection and maintenance of options and context-specific sequences of behavior directed towards particular goals. In line with this, it has been suggested that poorer monitoring of behavior by the aMCC (reflected as increased activity in the aMCC during error-related activity in a Go/NoGo task) may increase the effort required to inhibit behaviors. Highly relevant to our results, one previous study demonstrated that a network involving the aMCC and the SMA is critically involved in effort-based decision-making and the integration of motor costs into reward evaluation. More importantly, the same
study found that activity in the SMA was stronger in participants who tried to more activity avoid higher efforts. It is plausible therefore that those participants in our study who increased their time spent sedentary were engaging in effort avoidance and/or poor behavioral monitoring, which is reflected as an increase in FC between the aMCC (involved in decision making where motor costs are evaluated) and the SMA (has higher activity during effort avoidance). Our aMCC seed result also extended into the primary motor cortex (MI) as well (Fig. 3A). While voluntary movements and internally-selected actions are more traditionally associated with the SMA and aMCC, the TPJ and aMCC are functionally connected at rest and across multiple tasks, of contextual knowledge about salient stimuli. In accordance, the dAI has been suggested to be involved in other regions more associated with executive control and therefore may be more involved in the integration to salient stimuli. Indeed, the left medial portions of the cluster mapped onto the ventral attention network (VAN), a network involved in both attention and external awareness. Further, activity in the left motor cortex has been shown to increase as the subjective value of effortful rewards increases.

Higher FC between the r-dAI and a cluster overlapping the left temporoparietal and temporooccipital junction (regions covering the superior middle temporal gyrus and the inferior angular gyrus and lateral occipital gyrus) was also predictive of increases in time spent sedentary. The r-dAI has been proposed to provide an early cognitive control response and when mapping this result to a large functional network parcellation, both the r-dAI and portions of this cluster (those in the temporoparietal junction (TPJ)) map onto a broad, bilateral VAN/salience network. Indeed, group level connectivity of the r-dAI ROI (Fig. 1D) shows positive FC with salience/VAN regions and is anticorrelated with the default mode network (a hallmark sign of the VAN). The VAN is said to be involved in re-direction of attention to behaviorally relevant stimuli and is implicated in more external awareness than the synonymous salience network. Previous research using FC have shown the VAN to be predominantly lateralized to the right hemisphere, nevertheless, bilateral TPJ has been found during mental effort and is likely involved in an attentional network linking behavioral responses to salient stimuli. Indeed, the left medial portions of the cluster mapped onto the ventral attention network (VAN), a network involved in both attention and external awareness. Further, activity in the left motor cortex has been shown to increase as the subjective value of effortful rewards increases.

In this study, we show that individual differences in the baseline FC of multiple brain regions previously implicated in effort avoidance and effort-based decision making predict future change in sedentary time in low-active older adults participating in a 6-month walking intervention. Leveraging mechanistic predictors of future sedentary behaviors will potentially lead to targeted interventions that result in sustained behavioral change.

**Data availability**

All data will be provided upon reasonable request to the corresponding author, without reservation.

**Code availability**

Code used in this manuscript includes R syntax for statistical analyses and will be shared upon request to the corresponding author.

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T.P.M. Conceptualization, design, analysis, interpretation of data, manuscript writing, A.K. analysis, interpretation of data, substantial revision, S.A.A. analysis, interpretation of data, M.G. design, substantial revision, A.N.C. analysis, interpretation of data, A.B. data acquisition, substantial revision, N.G. data acquisition, substantial revision, J.F. data acquisition, substantial revision, E.S. data acquisition, substantial revision, S.W.G. analysis, interpretation of data, C.H. design, substantial revision, E.M. Conceptualization, design, substantial revision, A.F.K. Conceptualization, design, analysis, interpretation of data, substantial revision.

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Competing interests
The authors declare no competing interests.
