Augmented expression of Ki-67 is correlated with clinicopathological characteristics and prognosis for lung cancer patients: an up-dated systematic review and meta-analysis with 108 studies and 14,732 patients

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Abstract

Background: Lung cancer ranks as the leading cause of cancer-related deaths worldwide and we performed this meta-analysis to investigate eligible studies and determine the prognostic effect of Ki-67.

Methods: In total, 108 studies in 95 articles with 14,732 patients were found to be eligible, of which 96 studies reported on overall survival (OS) and 19 studies reported on disease-free survival (DFS) with relation to Ki-67 expression in lung cancer patients.

Results: The pooled hazard ratio (HR) indicated that a high Ki-67 level could be a valuable prognostic factor for lung cancer (HR = 1.122 for OS, \( P < 0.001 \) and HR = 1.894 for DFS, \( P < 0.001 \)). Subsequently, the results revealed that a high Ki-67 level was significantly associated with clinical parameters of lung cancer including age (odds ratio, OR = 1.246 for older patients, \( P = 0.018 \)), gender (OR = 1.874 for males, \( P < 0.001 \)) and smoking status (OR = 3.087 for smokers, \( P < 0.001 \)). Additionally, significant positive correlations were found between Ki-67 overexpression and poorer differentiation (OR = 1.993, \( P = 0.003 \)), larger tumor size (OR = 1.436, \( P = 0.003 \)), and higher pathologic stages (OR = 1.867 for III-IV, \( P < 0.001 \)). Furthermore, high expression of Ki-67 was found to be a valuable predictive factor for lymph node metastasis positive (OR = 1.653, \( P < 0.001 \)) and advanced TNM stages (OR = 1.497 for stage III-IV, \( P = 0.024 \)). Finally, no publication bias was detected in any of the analyses.

Conclusions: This study highlights that the high expression of Ki-67 is clinically relevant in terms of the prognostic and clinicopathological characteristics for lung cancer. Nevertheless, more prospective well-designed studies are warranted to validate these findings.

Keywords: Ki-67, Lung cancer, Meta-analysis, Prognosis, Clinicopathological characteristics
Background
Lung cancer is the most frequent diagnosed malignant neoplasms, and it was the first cause of cancer death in 2016 globally [1]. In the United States, lung cancer accounted for 27% of all cancer deaths in 2016. Non-small cell lung cancer (NSCLC), accounting for over 80% of all lung cancers, is the major cause of death worldwide [2]. Although the treatment of NSCLC patients includes surgery, radiotherapy and chemotherapy, the progress in lung cancer treatment is still slow, for which the 5-year relative survival is currently 18%. Diagnosis at an advanced stage is the major reason for this low survival rate [3]. Several prognostic factors were well characterized in lung cancer including sex, age, loss of weight, TNM stage, LDH, neutrophilia, haemoglobin as well as serum calcium [4]. Importantly, The IASLC Lung Cancer Staging Project in 2015 also carried out the 8th edition of the anatomic classification of lung cancer, which redefined the tumor-size cut-points in TNM stage for the lung cancer patients. The prognostic value of reclassification of tumor size were confirmed in 70,967 non-small-cell lung cancer patients from 1999 to 2010 [5]. To improve the survival of lung cancer patients, the choice of targeted treatments is increasingly being based on oncogenic drivers including ALK rearrangements, KRAS and epidermal growth factor receptor (EGFR) mutations [6, 7]. Additionally, BRAF mutations repre-
tended in lung cancer including sex, age, loss of weight, TNM stage, LDH, neutrophilia, haemoglobin as well as serum calcium [4]. Importantly, The IASLC Lung Cancer Staging Project in 2015 also carried out the 8th edition of the anatomic classification of lung cancer, which redefined the tumor-size cut-points in TNM stage for the lung cancer patients. The prognostic value of reclassification of tumor size were confirmed in 70,967 non-small-cell lung cancer patients from 1999 to 2010 [5]. To improve the survival of lung cancer patients, the choice of targeted treatments is increasingly being based on oncogenic drivers including ALK rearrangements, KRAS and epidermal growth factor receptor (EGFR) mutations [6, 7]. Additionally, BRAF mutations repre-
represented new therapeutic targets for lung cancer [8]. Likewise, several driver biomarkers also shed new light on the target treatment for lung cancer patients such as Her2 [9], NUT [10], DDR2 [11], FGFR1 [12], and PTEN [13]. Furthermore, several new molecular targets been highlighted in lung cancer, including ROS1 fusions [14], NTRK1 fusions [15] and exon 14 skipping mutations [16]. Recently, the checkpoint inhibitors targeting programmed death protein 1 (PD-1) have shown for durable clinical responses in NSCLC patients with advanced stage [17]. The immunomodulatory monoclonal antibodies against cytotoxic T-lymphocyte-associated antigen 4 (CTLA-4) also present promising results for the treatment of advanced stages lung cancer patients. However, due to the complex molecular mechanism of lung cancer, the identification of biomarkers in a large proportion of lung cancer patients is still required for a deeper understanding of the underlying epigenetic heterogeneity, and this will benefit the discovery of targeted therapies against lung cancer.

Ki-67, encoded by the MKI67 gene, is expressed throughout the cell cycle in proliferating but absent in in quiescent (G0) cells [18]. Ki-67 appears in the middle or late G1 stage, and then its expression increases through the S and G2 stage until it reaches a peak during the M stage [19]. The high expression of Ki-67 may contribute to aggressive and infiltrative growth of lung squamous carcinoma (SQC), cervical SQC and laryngeal SQC [20–22]. In addition, overexpression of Ki-67 has been positively associated with lymph node metastasis in gastric carcinoma and breast cancer [23, 24]. Ki-67 expression has been reported to be associated with a poor outcome in many malignancies including prostate, bladder and breast cancer [25–29].

Although two meta-analyses have previously reported that Ki-67 could be a possible indicator of short-term survival in lung cancer patients [30, 31], studies on a larger number of lung cancer patients and more reliable evidence are still needed to confirm the prognostic and clinicopathological role of Ki-67 for patients with lung cancer. Thus, our investigation aimed to evaluate the prognostic value and clinicopathological significance of Ki-67 in lung cancer patients via the review of previously published articles.

Results
Study selection and characteristics
As shown in Fig. 1, two investigators read the full text and considered 108 studies in 95 articles [18, 25, 32–124] consisting of 14,732 patients as applicable. The baseline characteristics of the included articles are indicated in Table 1. Ninety-six studies included data regarding OS, and nineteen studies included data regarding DFS. The number of cohorts of each study ranged from 32 to 778. In terms of the study region, 62 studies were from Asia, 30 were from Europe, and 15 were from America. In total, 104 studies consisted of 14,596 NSCLC patients: of these, 28 studies reported on ADC patients, and 10 studies reported on SQC patients. All the studies detected Ki-67 reactivity using IHC.

Prognostic value of Ki-67 for survival outcome in lung cancer patients
In total, 95 studies with 13,678 lung cancer patients investigated the impact of Ki-67 expression on OS (Table 2). The pooled HR of the total population for OS was 1.122 (95%CIs: 1.089–1.156, Z = 7.56, P < 0.001; I² = 78.20%, P < 0.001, Figs. 2, 3 and 4), showing that a high Ki-67 level indicates worse outcome for lung cancer patients. Furthermore, the correlation of high Ki-67 expression with DFS in 3127 lung cancer patients was then analyzed (Table 3). For the total study population, worse DFS (HR = 1.894, 95%CIs: 1.456–2.463, Z = 4.76, P < 0.001, Fig. 5) was observed among patients with high expression of Ki-67, while the heterogeneity using the random effects model was obvious (I² = 78.30%, P < 0.001). To investigate the source of heterogeneity, subgroup analyses of publication year, region, histological type, sample size, cut-off value of Ki-67 and estimated method for HR determination were performed. From the
subgroup analysis of OS, no heterogeneity was found in the small cell lung cancer (SCLC) group ($I^2 = 22.30\%, P = 0.277$). Next, a reduction in heterogeneity was observed after performing subgroup analysis of DFS according to the study region, especially in the studies from America ($I^2 = 8.40\%, P = 0.351$) and Asia ($I^2 = 25.60\%, P = 0.179$). As indicated in the subgroup of cut-off value, there was a low degree of heterogeneity in both Ki-67 low expression ($I^2 = 25.30\%, P = 0.196$) and high expression groups ($I^2 = 19.40\%, P = 0.287$). Furthermore, we also performed meta-regression analysis to explore the original of the heterogeneity in the studies. Consistent with the subgroup analysis, the results revealed that the regions and cut-off values might be the potential bias for the heterogeneity ($P = 0.017$ and $P = 0.022$, respectively). Altogether, we concluded that the different regions and inconsistent cut-off values might have contributed to the heterogeneity in the results of analyses for DFS. The regression also revealed that the heterogeneity originated from regions and inconsistent cut-off values (Table 4).

**The correlation of Ki-67 expression and clinicopathological features in lung cancer patients**

An association of Ki-67 expression with age in 2506 lung cancer patients was identified using the fixed effects model in 19 studies, and higher Ki-67 expression was found to be more common in older patients (OR = 1.246, 95%CIs: 1.039–1.494; $Z = 2.37$, $P = 0.018$, $I^2 = 0.00\%$, $P = 0.967$, Table 5 and Additional file 1: Figure S1). Subsequently, the results revealed significant differences in Ki-67 level between male and female (OR = 1.874, 95%CIs: 1.385–2.535; $Z = 4.07$, $P < 0.001$, $I^2 = 69.70\%$, $P < 0.0001$, Additional file 1: Figure S1B). Meta-analysis of 15 studies including 2152 lung cancer patients revealed a positive association between high Ki-67 level and smoking history (OR = 3.087, 95%CIs: 2.504–3.806, $Z = 10.56$, $P < 0.001$; $I^2 = 39.40\%$, $P = 0.064$, Additional file 1: Figure S1C). According to the histological type, a pooled OR of 0.397 (95%CIs: 0.236–0.667) indicated that Ki-67 expression was significantly higher in ADC compared with that in SQC ($Z = 3.49$, $P < 0.001$; $I^2 = 81.20\%$, $P < 0.001$, Additional file 2: Figure S2A). Next, tumor differentiation was considered. The results from 11 studies enrolling 1731 lung cancer patients showed that an elevated Ki-67 level was associated with poor differentiation, with a pooled OR of 1.993 (95%CIs:1.262–3.146, $Z = 2.96$, $P = 0.003$; $I^2 = 66.30\%$, $P = 0.001$, Additional file 2: Figure S2B). A total of 13 studies with 1851 individuals were analyzed in this meta-analysis, and the results showed that a higher Ki-67 level was positively associated with the pathologic stage III/IV with a low degree of heterogeneity (OR = 1.867, 95%CIs:1.498–2.327, $Z = 5.56$, $P < 0.001$; $I^2 = 23.1\%$, $P = 0.210$, Additional file 2: Figure S2C). A trend toward positive correlation was found between a high Ki-67 level and larger tumor size in 12 studies based on 1707 lung cancer patients, with a pooled OR of 1.436 (95%CIs:1.127–1.290, $Z = 2.93$, $P = 0.003$; $I^2 = 0.00\%$, $P = 0.876$, Additional file 3: Figure S3A). Twenty-three studies comprising 2994 cases were used for meta-analysis of Ki-67 expression and lymph node metastasis, and the pooled OR indicated that a high Ki-67 level was significantly correlated with lymph node metastasis positive (OR = 1.653, 95%CIs: 1.285–2.127, $Z = 3.91$, $P < 0.0001$; $I^2 = 46.70\%$, $P = 0.008$, Additional file 3: Figure S3B). The association of Ki-67 expression and TNM stage
| Study (author/year) | Region          | Tumor stage | Histological type | Patients | Sample size | Cutoff value | Sample type     | Assay          | NOS score | Extract method | Survival  |
|---------------------|-----------------|-------------|-------------------|----------|-------------|--------------|----------------|----------------|-----------|----------------|-----------|
| Scagliotti 1993 [91]| Italy           | I-IIIA      | NSCLC             | 111      | Large       | 25%          | Tumor tissue   | IHC            | 7         | Survival curve | OS        |
| Pence 1993 [120]    | USA             | I-IV        | NSCLC             | 61       | Small       | 4%           | Tumor tissue   | IHC            | 7         | Survival curve | OS        |
| Fontanini 1996 [46] | Italy           | NA          | NSCLC             | 70       | Small       | 30%          | Tumor tissue   | IHC            | 7         | Survival curve | OS        |
| Bohm 1996 [37]      | Germany         | NA          | SCLC              | 32       | Small       | 27%          | Tumor tissue   | IHC            | 7         | Original data  | OS        |
| Harpole 1996 [50]   | USA             | I           | NSCLC             | 275      | Large       | 7%           | Tumor tissue   | IHC            | 9         | Multivariate   | OS        |
| Pujoll 1996 [88]    | France          | I-IV        | LC                | 97       | Small       | NA           | Tumor tissue   | IHC            | 7         | Survival curve | OS        |
| Mehdi 1999 [119]    | USA             | I-IV        | NSCLC             | 203      | Large       | 25%          | Tumor tissue   | IHC            | 9         | Multivariate   | OS        |
| Demarchi 1999 [41]  | Brazil          | I-III       | ADC               | 64       | Small       | 22%          | Tumor tissue   | IHC            | 7         | Original data  | OS        |
| Dingemans 1999 [42] | Netherland      | I-III       | SCLC              | 93       | Small       | 30%          | Tumor tissue   | IHC            | 7         | Survival curve | OS        |
| Wang 1999 [99]      | China           | NA          | NSCLC             | 85       | Small       | 30%          | Tumor tissue   | IHC            | 7         | Survival curve | OS        |
| Shiba 2000 [92]     | Japan           | I-III       | NSCLC             | 95       | Small       | 20%          | Tumor tissue   | IHC            | 7         | Univariate     | OS        |
| Hommura 2000 [54]   | Japan           | I-II        | SQC               | 91       | Small       | 30%          | Tumor tissue   | IHC            | 9         | Multivariate   | OS        |
| Hommura 2000 [54]   | Japan           | I-II        | NSCLC             | 124      | Large       | 30%          | Tumor tissue   | IHC            | 9         | Multivariate   | OS        |
| Nguyen 2000 [80]    | Czech Republic  | I-IV        | NSCLC             | 89       | Small       | 30%          | Tumor tissue   | IHC            | 7         | Survival curve | OS        |
| Puglisi 2001 [87]   | Italy           | I-III       | NSCLC             | 81       | Small       | 30%          | Tumor tissue   | IHC            | 9         | Multivariate   | OS        |
| Hayashi 2001 [52]   | Japan           | I-IV        | NSCLC             | 98       | Small       | 13%          | Tumor tissue   | IHC            | 9         | Multivariate   | OS        |
| Pelosi 2001 [84]    | Italy           | I           | SQC               | 119      | Large       | NA           | Tumor tissue   | IHC            | 9         | Multivariate   | OS        |
| Ramnath 2001 [121]  | USA             | I-IV        | NSCLC             | 160      | Large       | 24%          | Tumor tissue   | IHC            | 7         | Univariate     | OS        |
| Ramnath 2001 [121]  | USA             | I-IV        | NSCLC             | 41       | Small       | 50%          | Tumor tissue   | IHC            | 7         | Univariate     | OS        |
| Wang 2001 [101]     | China           | I-IV        | LC                | 166      | Large       | 18%          | Tumor tissue   | IHC            | 7         | Survival curve | OS        |
| Mojtahedzadeh 2002 [78] | Japan     | I-III       | ADC               | 141      | Large       | 10%          | Tumor tissue   | IHC            | 8         | Multivariate   | OS        |
| Minami 2002 [76]    | Japan           | I           | ADC               | 47       | Small       | 20%          | Tumor tissue   | IHC            | 8         | Multivariate   | OS        |
| Takahashi 2002 [94] | Japan           | I-IV        | NSCLC             | 62       | Small       | 25%          | Tumor tissue   | IHC            | 9         | Multivariate   | DFS       |
| Wakabayashi 2003 [97]| Japan           | I-IV        | NSCLC             | 140      | Large       | 13%          | Tumor tissue   | IHC            | 9         | Multivariate   | OS        |
| Pelosi 2003 [85]    | UK              | I           | ADC               | 96       | Small       | NA           | Tumor tissue   | IHC            | 9         | Multivariate   | OS        |
| Haga 2003 [49]      | Japan           | I           | ADC               | 58       | Small       | 10%          | Tumor tissue   | IHC            | 9         | Multivariate   | OS        |
| Hashimoto 2003 [51] | Japan           | I-III       | ADC               | 122      | Large       | 20%          | Tumor tissue   | IHC            | 8         | Multivariate   | OS        |
| Poleri 2003 [86]    | Argentina       | I           | NSCLC             | 50       | Small       | 67%          | Tumor tissue   | IHC            | 7         | Survival curve | DFS       |
| Matheus 2004 [75]   | Brazil          | I-III       | ADC               | 33       | Small       | 22%          | Tumor tissue   | IHC            | 7         | Original data  | OS        |
| Niemiec 2004 [81]   | Poland          | I-III       | SQC               | 78       | Small       | 28%          | Tumor tissue   | IHC            | 7         | Survival curve | OS        |
| Study (author/year) | Region | Tumor stage | Histological type | Patients | Sample size | Cutoff value | Sample type | Assay | NOS score | Extract method | Survival |
|---------------------|--------|-------------|-------------------|----------|-------------|--------------|-------------|-------|-----------|----------------|----------|
| Ahn 2004 [33] | Korea | II-III A | NSCLC | 65 | Small | 15% | Tumor tissue | IHC | 9 | Multivariate | OS |
| Huang 2005 [55] | Japan | I | NSCLC | 97 | Small | 25% | Tumor tissue | IHC | 8 | Multivariate | OS |
| Gasinska 2005 [48] | Poland | I-III | SQC | 81 | Small | 39% | Tumor tissue | IHC | 9 | Multivariate | OS |
| Wang 2005 [100] | China | I-III | NSCLC | 51 | Small | 5% | Tumor tissue | IHC | 7 | Survival curve | OS |
| Dong 2005 [43] | Japan | I-IV | ADC | 131 | Large | 18% | Tumor tissue | IHC | 9 | Multivariate | OS |
| Niemiec 2005 [81] | Poland | I-III | SQC | 78 | Large | 28% | Tumor tissue | IHC | 7 | Survival curve | OS |
| Huang 2005 [55] | Japan | II-III | NSCLC | 76 | Small | 25% | Tumor tissue | IHC | 8 | Multivariate | OS |
| Tsubochi 2006 [64] | Japan | I-III | NSCLC | 219 | Large | 20% | Tumor tissue | IHC | 9 | Multivariate | OS |
| Yang 2006 [110] | USA | I-III | NSCLC | 128 | Large | 25% | Tumor tissue | IHC | 9 | Multivariate | OS |
| Nozawa 2006 [82] | Japan | IV | ADC | 35 | Small | 40% | Tumor tissue | IHC | 7 | Survival curve | OS |
| Maddau 2006 [118] | Italy | II-III | NSCLC | 88 | Large | 25% | Tumor tissue | IHC | 7 | Univariate | OS |
| Maddau 2006 [118] | Italy | I | NSCLC | 92 | Large | 25% | Tumor tissue | IHC | 7 | Univariate | OS |
| Inoue 2007 [58] | Japan | I-III | ADC | 97 | Small | 5% | Tumor tissue | IHC | 9 | Multivariate | DFS |
| Mohamed 2007 [77] | Japan | I-IV | NSCLC | 61 | Small | 20% | Tumor tissue | IHC | 9 | Multivariate | OS |
| Zhou 2007 [116] | China | I-II | NSCLC | 70 | Small | NA | Tumor tissue | IHC | 6 | Multivariate | OS |
| Moreno 2007 [79] | Argentina | III | NSCLC | 32 | Small | 66% | Tumor tissue | IHC | 7 | Survival curve | OS |
| Yoo 2007 [112] | Korea | I-III | NSCLC | 219 | Large | 30% | Tumor tissue | IHC | 9 | Multivariate | OS |
| Fujioka 2008 [47] | Japan | I | ADC | 73 | Small | 14% | Tumor tissue | IHC | 9 | Multivariate | OS |
| Imai 2008 [57] | Japan | I | NSCLC | 248 | Large | 25% | Tumor tissue | IHC | 9 | Multivariate | OS |
| Woo 2008 [103] | Japan | Ia | ADC | 131 | Large | 10% | Tumor tissue | IHC | 8 | Univariate | DFS |
| Woo 2008 [103] | Japan | Ib | ADC | 59 | Small | 10% | Tumor tissue | IHC | 8 | Univariate | DFS |
| Saad 2008 [89] | USA | I-III | NSCLC | 54 | Small | 30% | Tumor tissue | IHC | 7 | Survival curve | OS |
| Kaira 2008 [124] | Japan | I-III | NSCLC | 321 | Large | 25% | Tumor tissue | IHC | 9 | Multivariate | OS |
| Ikeda 2008 [56] | Japan | I-III | NSCLC | 200 | Large | 5% | Tumor tissue | IHC | 7 | Univariate | OS and DFS |
| Anami 2009 [34] | Japan | I-IV | ADC | 139 | Large | 10% | Tumor tissue | IHC | 8 | Multivariate | OS |
| Yuan 2009 [113] | China | I-III | NSCLC | 140 | Large | 25% | Tumor tissue | IHC | 8 | Multivariate | OS |
| Kaira 2009 [62] | Japan | I | ADC | 139 | Large | 20% | Tumor tissue | IHC | 9 | Multivariate | OS |
| Erler 2010 [44] | USA | NA | SCLC | 68 | Small | 50% | Tumor tissue | IHC | 7 | Survival curve | OS |
| Filipits 2011 [45] | Austria | I-III | NSCLC | 778 | Large | NA | Tumor tissue | IHC | 9 | Multivariate | OS and DFS |
| Werynska 2011 [102] | Poland | I-IV | NSCLC | 145 | Large | 25% | Tumor tissue | IHC | 7 | Univariate | OS |
| Yamashita 2011 [109] | Japan | I | NSCLC | 44 | Small | 5% | Tumor tissue | IHC | 8 | Multivariate | DFS |
| Wu 2011 [106] | China | I-IV | NSCLC | 160 | Large | 10% | Tumor tissue | IHC | 8 | Multivariate | OS |
| Oka 2011 [83] | Japan | I-III | ADC | 183 | Large | 20% | Tumor tissue | IHC | 9 | Multivariate | DFS |
| Sterlacchi 2011 [122] | Austria | I-IV | NSCLC | 386 | Large | 3% | Tumor tissue | IHC | 9 | Multivariate | OS |
| Werynska 2011 [102] | Poland | NA | NSCLC | 145 | Large | 25% | Tumor tissue | IHC | 7 | Univariate | OS |
| Liu 2012 [71] | China | I-IV | NSCLC | 494 | Large | 50% | Tumor tissue | IHC | 9 | Multivariate | OS |
| Wang 2012 [98] | China | NA | SCLC | 42 | Small | 10% | Tumor tissue | IHC | 7 | Survival curve | OS |
Table 1 Characteristics of studies included into the meta-analysis (Continued)

| Study (author/year) | Region | Tumor stage | Histological type | Patients | Sample size | Cutoff value | Sample type | Assay | NOS score | Extract method | Survival value |
|---------------------|--------|-------------|-------------------|----------|-------------|--------------|-------------|-------|-----------|----------------|----------------|
| Liu 2012 [71]       | China  | I-IV        | ADC               | 97       | Small       | 10%          | Tumor tissue| IHC   | 7         | Survival curve | OS             |
| Wu 2012 [105]       | China  | I-IV        | ADC               | 309      | Large       | 50%          | Tumor tissue| IHC   | 9         | Multivariate OS | OS             |
| Salvi 2012 [90]     | Italy  | I-III       | NSCLC             | 81       | Small       | 15%          | Tumor tissue| IHC   | 9         | Multivariate OS | OS             |
| Yang 2012 [111]     | China  | I-III       | NSCLC             | 68       | Large       | 38%          | Tumor tissue| IHC   | 7         | Survival curve | OS             |
| Wu 2013 [104]       | China  | I-IV        | NSCLC             | 192      | Large       | 10%          | Tumor tissue| IHC   | 9         | Multivariate OS | OS and DFS     |
| Maki 2013 [74]      | Japan  | I           | ADC               | 105      | Large       | 15%          | Tumor tissue| IHC   | 9         | Multivariate DFS| OS             |
| Lei 2013 [69]       | China  | I-IV        | NSCLC             | 279      | Large       | 30%          | Tumor tissue| IHC   | 9         | Multivariate OS | OS             |
| Berghoff 2013 [36]  | Austria| I-IV        | NSCLC             | 230      | Large       | 40%          | Tumor tissue| IHC   | 9         | Multivariate OS | OS             |
| Ji 2013 [61]        | China  | I-III       | NSCLC             | 67       | Small       | 5%           | Tumor tissue| IHC   | 9         | Multivariate OS | OS             |
| Kobayakov 2013 [68] | USA    | I-III       | SQC               | 118      | Large       | 30%          | Tumor tissue| IHC   | 7         | Survival curve | OS             |
| Zu 2013 [117]       | China  | I-III       | ADC               | 96       | Small       | 25%          | Tumor tissue| IHC   | 7         | Survival curve | OS             |
| Liu 2013 [70]       | China  | I-III       | NSCLC             | 105      | Large       | 50%          | Tumor tissue| IHC   | 8         | Multivariate OS | OS             |
| Hokka 2013 [53]     | Japan  | I-IV        | ADC               | 125      | NA          | Tumor tissue | IHC   | 9         | Multivariate OS | OS             |
| Xue 2013 [73]       | China  | I-III       | NSCLC             | 83       | Small       | 50%          | Tumor tissue| IHC   | 8         | Multivariate OS | OS             |
| Zhong 2014 [115]    | China  | I-IV        | NSCLC             | 270      | Large       | 50%          | Tumor tissue| IHC   | 8         | Multivariate OS | OS             |
| Ahn 2014 [32]       | Korea  | I-III       | NSCLC             | 108      | Large       | 40%          | Tumor tissue| IHC   | 9         | Multivariate DFS| OS             |
| Shimizu 2014 [93]   | Japan  | I-III       | SQC               | 32       | Small       | 10%          | Tumor tissue| IHC   | 8         | Multivariate DFS| OS             |
| Shimizu 2014 [93]   | Japan  | I-III       | ADC               | 52       | Small       | 10%          | Tumor tissue| IHC   | 8         | Multivariate DFS| OS             |
| Kim 2014 [136]      | Korea  | I-IV        | ADC               | 122      | Large       | 10%          | Tumor tissue| IHC   | 7         | Univariate OS   | OS             |
| Tsoukalas 2014 [95] | Greece | I-IV        | NSCLC             | 112      | NA          | Tumor tissue | IHC   | 9         | Multivariate OS | OS             |
| Kawatsu 2014 [63]   | Japan  | I-IV        | NSCLC             | 183      | Large       | 10%          | Tumor tissue| IHC   | 8         | Multivariate OS | OS             |
| Corzani 2014 [39]   | Italy  | III          | NSCLC             | 50       | Large       | 50%          | Tumor tissue| IHC   | 8         | Multivariate OS | OS             |
| Warth 2014 [18]     | Germany| I-IV        | ADC               | 482      | Large       | 25%          | Tumor tissue| IHC   | 7         | Univariate OS   | OS and DFS     |
| Warth 2014 [18]     | Germany| I-IV        | SQC               | 233      | Large       | 50%          | Tumor tissue| IHC   | 8         | Multivariate OS | OS             |
| Tabata 2014 [25]    | Japan  | I-IV        | NSCLC             | 74       | Small       | 10%          | Tumor tissue| IHC   | 9         | Multivariate OS | OS             |
| Xu 2014 [108]       | China  | I-IV        | ADC               | 80       | Small       | 5%           | Tumor tissue| IHC   | 7         | Survival curve  | OS             |
| Liu 2014 [70]       | China  | I-IV        | NSCLC             | 96       | Small       | 30%          | Tumor tissue| IHC   | 7         | Survival curve  | OS             |
| Ji 2014 [60]        | China  | I-III       | NSCLC             | 83       | Small       | NA           | Tumor tissue | IHC   | 8         | Multivariate OS | OS             |
| Shimizu 2014 [93]   | Japan  | I-IV        | SQC               | 32       | Large       | 10%          | Tumor tissue| IHC   | 9         | Multivariate OS | OS             |
| Shimizu 2014 [93]   | Japan  | I-IV        | ADC               | 52       | Large       | 10%          | Tumor tissue| IHC   | 9         | Multivariate OS | OS             |
| Kobierzycki 2014 [67]| Poland| I-IV        | NSCLC             | 218      | Large       | 25%          | Tumor tissue| IHC   | 6         | Univariate OS   | OS             |
| Xu 2014 [107]       | China  | I-III       | NSCLC             | 114      | Large       | 50%          | Tumor tissue| IHC   | 7         | Univariate OS   | OS             |
| Zhang 2015 [114]    | China  | I-IV        | ADC               | 616      | Large       | NA           | Tumor tissue| IHC   | 8         | Multivariate OS | OS             |
| Gobbo 2015 [40]     | Italy  | NA           | NSCLC             | 383      | Large       | 20%          | Tissue microarray | IHC   | 7         | Univariate OS   | OS             |
| Stewart 2015 [123]  | USA    | II-III      | NSCLC             | 230      | Large       | NA           | Tumor tissue| IHC   | 7         | Univariate DFS   | OS             |
| Vignonroux 2015 [96]| France | I-IV        | NSCLC             | 190      | Large       | 40%          | Tumor tissue| IHC   | 7         | Survival curve  | OS             |
was then incorporated into the meta-analysis. Eight studies with 736 patients showed a trend for correlation between Ki-67 overexpression and advanced TNM stages, with a pooled OR of 1.50 (95%CIs:1.053–2.126, Z = 2.25, P = 0.024; I² = 36.90%, P = 0.134, Additional file3: Figure S3C). In the meta-analysis, no association between Ki-67 and tumor stage was observed in lung cancer patients (OR = 1.287, 95%CIs:0.882–1.877, Z = 1.31, P = 0.191;I² = 55.30%, P = 0.013). Additionally, analysis of four selected studies using the random effects model did not reveal any significance for the association between Ki-67 expression and metastasis (OR = 2.609, 95%CIs: 0.667–10.204, Z = 1.38, P = 0.168) or invasion (OR = 0.993, 95%CIs: 0.511–1.930, Z = 0.02, P = 0.984; I² = 14.20%, P = 0.312).

To identify potential publication bias, Begg’s test and funnel plots were used. No publication bias was found in the analysis for OS (p = 0.444, Fig. 6a) and DFS (P = 0.246, Fig. 6b). Moreover, there was no publication bias for the association between Ki-67 expression and metastasis (OR = 2.609, 95%CIs: 0.667–10.204, Z = 1.38, P = 0.168) or invasion (OR = 0.993, 95%CIs: 0.511–1.930, Z = 0.02, P = 0.984; I² = 14.20%, P = 0.312).

**Table 1** Characteristics of studies included into the meta-analysis (Continued)

| Study(author/year) | Region | Tumor stage | Histological type | Patients | Sample size | Cutoff value | Sample type | NOS score | Extract method | Survival |
|-------------------|--------|-------------|-------------------|----------|-------------|--------------|-------------|----------|----------------|----------|
| Jethon 2015 [59]  | Poland | I-IV        | SQC               | 89       | Small       | 25%          | Tumor tissue | IHC       | Univariate     | OS       |
| Jethon 2015 [59]  | Poland | I-IV        | ADC               | 98       | Small       | 25%          | Tumor tissue | IHC       | Univariate     | OS       |
| Apostolova 2016   | Germany| I-IV        | NSCLC             | 83       | Small       | 75%          | Tumor tissue | IHC       | Multivariate   | OS       |
| Cardona 2016 [38] | USA    | NA          | NSCLC             | 144      | Large       | 30%          | Tumor tissue | IHC       | Original data  | OS       |

**Table 2** Summarized HRs of overall and subgroup analyses for OS

| Stratified analysis | Study(N) | HR (95%CI) | z    | P    | Heterogeneity | P    | Estimated method |
|-------------------|----------|------------|------|------|---------------|------|------------------|
| OS                | 95       | 1.122(1.089–1.156) | 7.56 | < 0.001 | 78.20%        | < 0.001 | Random-effect     |
| Publication year   |          |            |      |      |               |      |                  |
| Early year(~ 2007) | 44       | 1.307(1.212–1.408) | 7.01 | < 0.001 | 72.50%        | < 0.001 | Random-effect     |
| Later year(2007~ 2016) | 51   | 1.101(1.063–1.142) | 5.28 | < 0.001 | 81.20%        | < 0.001 | Random-effect     |
| Region             |          |            |      |      |               |      |                  |
| Europe             | 30       | 1.021(1.001–1.042) | 2.05 | 0.041   | 71.30%        | < 0.001 | Random-effect     |
| America            | 13       | 1.671(1.266–2.205) | 3.63 | < 0.001 | 66.60%        | < 0.001 | Random-effect     |
| Asia               | 52       | 1.821(1.623–2.043) | 10.21| < 0.001 | 78.20%        | < 0.001 | Random-effect     |
| Histological type  |          |            |      |      |               |      |                  |
| SCLC               | 4        | 1.023(1.004–1.042) | 3.28 | 0.001  | 22.30%        | 0.277 | Fixed-effect      |
| NSCLC              | 89       | 1.113(1.081–1.147) | 7.11 | < 0.001 | 78.70%        | < 0.001 | Random-effect     |
| ADC                | 22       | 1.219(1.113–1.338) | 4.26 | < 0.001 | 76.20%        | < 0.001 | Random-effect     |
| SQC                | 9        | 1.115(0.806–1.542) | 0.66 | 0.512   | 74.40%        | < 0.001 | Random-effect     |
| Sample size        |          |            |      |      |               |      |                  |
| < 100              | 45       | 1.486(1.340–1.649) | 7.71 | < 0.001 | 73.10%        | < 0.001 | Random-effect     |
| > 100              | 50       | 1.083(1.049–1.119) | 4.90 | < 0.001 | 81.50%        | < 0.001 | Random-effect     |
| Cutoff value       |          |            |      |      |               |      |                  |
| L(< 20%)           | 34       | 1.962(1.622–2.373) | 6.95 | < 0.001 | 74.50%        | < 0.001 | Random-effect     |
| H(≥20%)            | 52       | 1.144(1.094–1.197) | 5.91 | < 0.001 | 78.90%        | < 0.001 | Random-effect     |
| Estimated method   |          |            |      |      |               |      |                  |
| Original data      | 4        | 2.043(0.868–4.808) | 1.64 | 0.102  | 83.90%        | < 0.001 | Random-effect     |
| Survival curve     | 23       | 1.629(1.368–1.940) | 5.47 | < 0.001 | 76.30%        | < 0.001 | Random-effect     |
| HR(univariate)     | 16       | 1.511(1.236–1.847) | 4.02 | < 0.001 | 56.10%        | 0.004  | Random-effect     |
| HR(multivariate)   | 53       | 1.108(1.063–1.155) | 4.87 | < 0.001 | 75.30%        | < 0.001 | Random-effect     |
among any of the analyses used to correlate Ki-67 expression and clinicopathological characteristics (all $P > 0.05$, Table 5, Additional file 4: Figure S4, Additional file 5: Figure S5, Additional file 6: Figure S6 and Additional file 7: Figure S7).

**Discussion**

As previously mentioned, there are two meta-analyses showing that high expression of Ki-67 predicts worse prognosis in lung cancer patients [30] and early-stage NSCLCs [31]. Nevertheless, there is no consensus regarding the clinicopathological significance of Ki-67 in lung cancer patients. Martin et al. performed a meta-analysis on 37 studies to evaluate the prognostic value of Ki-67 in 3983 lung cancer patients in 2004 [30].

Lacking sufficient information for other subtypes of lung cancer and Asian patients, the results from the above-mentioned meta-analysis were not convincing. Our meta-analysis includes 108 studies with 14,831 lung cancer patients comprising NSCLC and SCLC cases and thus provides more reliable evidence. Additionally, we also restricted the number of patients in each study to greater than 30 to exclude low-quality studies. To strengthen the evidence, we estimated not only OS data but also DFS to determine the prognostic role of Ki-67 in lung cancer patients. Moreover, multivariate analyses of OS and DFS were performed, and the HR of OS was 1.108(95% CIs: 1.063–1.155), and that for DFS was 1.892(95% CIs: 1.328–2.698), indicating that Ki-67 is an independent prognostic marker for lung cancer.
Compared to the previous meta-analysis, our meta-analysis included results on all subtypes of lung cancer and represented broader ethnicity; in addition, subgroups were classified according to region, cut-off value, number of patients and histological type. With the inclusion of high-quality studies and a larger number of patients, the results derived from our study are more convincing.

Ki-67 is present in the active phases of the cell cycle (G1, S, and G2), as well as during mitosis, but it is not expressed in the G0 phase. Thus, it has become an excellent operational marker for the estimation of the proportion of proliferative cells in a given cell population [125]. Our study demonstrated that Ki-67 expression was lower in ADC compared with that in SQC, suggesting it is a useful biomarker for distinguishing ADC from SQC. Consisted with our result, Ki-67 was also revealed to be higher in ADC than in SQC in the same stage, due to the different tumor biology of histological subtypes in NSCLC [49, 64]. The high-grade Ki-67 was proved to be significantly correlated with a more aggressive tumor infiltration patterns in lung SQC, indicating the strong association between tumor invasiveness and cell proliferation [20]. It has been reported in the literature that inverted papillomas with high levels of Ki-67 also include squamous cell carcinomas. Regarding the SCLC, Vasudha Murlidhar et al. recently carried out a result that Ki-67 could contribute to the early detection of metastasis in circulating lung cancer cells. Ki-67 was also demonstrated as a potential diagnostic
factor for histopathological definition of SCLC [126]. Contrary to our study, previous study revealed that maternal cigarette smoking could dramatically decrease the expression of Ki-67 in cytotrophoblasts [127]. Interestingly, further study also found that Ki-67 was lower expressed in smokers and smokers with COPD compared to the non-smokers. The authors hypothesized that the permanent cellular damage might play a crucial role in the destruction of bronchiolar tissue [128]. Nevertheless, the mechanisms that govern how Ki-67 expression contribute to the tumorigenesis and progression of lung cancer remain to be unveiled.

The authors suspected that Ki-67 might affect cyclin-dependent kinase1 (CDK1), leading to the entry of inverted papilloma cells into the active phase of cell cycle(G1) and resulting in malignant transformation [129]. Interestingly, our study found that Ki-67 expression in male patients was significantly higher comparing with that in female patients. Previous reports have suggested that testosterone can promote the growth of cancer cells that express androgen receptors, which negatively regulates the Ki-67 level in lung cancer patients [130, 131]. Many previous studies have also revealed that Ki-67 is significantly associated with histopathologic parameters in other tumor because of the correlation between proliferation and those parameters [132–134]. It was found that p53 regulated the p53-and Sp1-dependent pathways, leading to the inhibition of Ki-67 promoter [135, 136]. A recent study confirmed that there was a correlation between the specific Ki-67
splice variants and the progression through the cell cycle in cancer cells. Ki-67 might be involved in a putative extranuclear elimination pathway transported to the Golgi apparatus [137]. Based on these results, we concluded that Ki-67 serves as a valuable indicator for the aggressiveness and prognosis of lung cancer.

Heterogeneity was significant in this meta-analysis. To eliminate the heterogeneity, subgroup analyses according to region, cut-off value, number of patients and histological type were carried out using random effects models. As a result, we revealed that the source of heterogeneity originated from the publication region and the cutoff value by the meta-regression analysis. Our meta-analysis was limited to publications in English or Chinese; nevertheless, the researchers typically tend to publish studies with negative results in local journals and in the native language of the study region. In addition, although detailed exclusion criteria were established to avoid duplication, our meta-analysis was not able to avoid the same patient cohorts in different publications. Other methodological factors might also affect heterogeneity, such as the antibody and cut-off value used in the study. Although anti-MIB-1 antibody is the most frequently used antibody in studies, most of the included studies stratified high and low levels of Ki-67 using a median value varying from 3 to 75%, which might have influenced the results. Several studies used the Ki-67 cut-off less than 10% to assess the prognostic impact of Ki-67 after surgical resection with curative intent in early-stages lung cancer patients. Most of the included studies used the median value of Ki-67 index as cut-off value, which could divide the patient into equally the group but did not reflect the clinical relevant use. A cut-off value of Ki-67 maximizing the hazard ratio across the groups could be used for the clinical management for the diagnosis and prognosis of lung cancers. More importantly, Multiple clinical laboratories have reported the Ki-67 cutoff values between 10 and 14% could be recommended as the gold standard identify the high risk of the survival outcome in cancers [138–141].

### Table 3 Summarized HRs of overall and subgroup analyses for DFS

| Stratified analysis | Study(N) | HR       | Z     | P     | Heterogeneity | Estimated method |
|--------------------|----------|----------|-------|-------|---------------|-----------------|
|                    |          |          |       |       |               |                 |
| DFS                | 21       | 1.894(1.456–2.463) | 4.76  | <0.001 | 78.30%        | Random-effect   |
|                    |          |          |       |       |               |                 |
| Subgroup analysis for DFS |          |          |       |       |               |                 |
| Publication year   |          |          |       |       |               |                 |
| Early year(~ 2007) | 6        | 1.428(0.992–2.055) | 1.92  | 0.055 | 62.10%        | Random-effect   |
| Later year(2007~2016) | 15       | 2.237(1.54–3.249) | 4.23  | <0.001 | 72.00%        | Random-effect   |
| Region             |          |          |       |       |               |                 |
| Europe             | 3        | 1.023(1.005–1.041) | 2.51  | 0.012 | 53.40%        | Fixed-effect    |
| America            | 4        | 1.559(1.155–2.105) | 2.9   | 0.004 | 8.40%         | Fixed-effect    |
| Asia               | 14       | 2.673(2.096–3.409) | 7.92  | <0.001 | 25.60%        | Fixed-effect    |
| Histological type  |          |          |       |       |               |                 |
| SCLC               | 21       | 1.894(1.456–2.463) | 4.76  | <0.001 | 78.30%        | Random-effect   |
| NSCLC              | 9        | 3.186(1.797–5.650) | 3.96  | <0.001 | 62.10%        | Random-effect   |
| ADC                | 2        | 1.022(1.004–1.04)  | 2.42  | 0.015 | 0.00%         | Fixed-effect    |
| SQC                | 2        | 1.022(1.004–1.04)  | 2.42  | 0.015 | 0.00%         | Fixed-effect    |
| Sample size        |          |          |       |       |               |                 |
| < 100              | 7        | 2.455(1.392–4.330) | 3.10  | <0.001 | 20.30%        | Fixed-effect    |
| > 100              | 14       | 1.770(1.340–2.338) | 4.02  | <0.001 | 82.80%        | Fixed-effect    |
| Cutoff value       |          |          |       |       |               |                 |
| L(< 20%)           | 12       | 2.783(2.141–3.619) | 7.64  | <0.001 | 25.30%        | Fixed-effect    |
| H(≥ 20%)           | 6        | 1.514(1.243–1.844) | 4.12  | <0.001 | 19.40%        | Fixed-effect    |
| Estimated method   |          |          |       |       |               |                 |
| Survival curve     | 2        | 1.595(1.053–2.416) | 2.21  | 0.027 | 52.60%        | Fixed-effect    |
| HR(univariate)     | 6        | 2.126(1.156–3.909) | 2.43  | 0.015 | 67.40%        | Random-effect   |
| HR(multivariate)   | 13       | 1.892(1.328–2.698) | 3.53  | <0.001 | 79.90%        | Random-effect   |

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Additionally, microarrays were used in several studies, for which the sensitivity of assessment of Ki-67 expression is generally poor. Another possibility of bias may be related to the method of extrapolating the HR; the HR extracted from survival curves was less reliable than direct analysis of variance. Additionally, the subgroup analysis for different stages because most of the included studies recruited the lung cancer patients within three or more tumor stages. Thus, we could not classify the patients into the early stages and advanced stages for the subgroup analysis.

**Conclusion**

In conclusion, our meta-analysis demonstrated that high expression of Ki-67 is associated with worse prognosis and disease progression in lung cancer patients. Ki-67 can be an independent biological marker for predicting the prognosis of lung cancer patients. Subsequent studies are required to investigate the prognoses and clinical characteristics of lung cancer patients to confirm our findings.

**Materials and methods**

**Literature search and selection**

The databases that we searched included PubMed, Web of Science, EMBASE, and Chinese datasets (WanFang, China National Knowledge Infrastructure and Chinese VIP) until June 1, 2017. The key words identifying the articles were as follows: (Ki-67 OR Ki67 OR MIB-1 OR “proliferative index” OR “proliferative activity” OR...
mitotic index” OR “labeling index” OR “mitotic count” OR “proliferative marker” OR “mitotic figure” OR “mitotic activity”) AND (Cancer OR carcinoma OR adenocarcinoma OR tumour OR tumor OR malignanc* OR neoplas*) AND (Lung OR pulmonary OR respiratory OR respiration OR aspiration OR bronchi OR bronchioles OR alveoli OR pneumocytes OR “air way”).

**Selection criteria**
Publications were included if they met the following inclusion criteria: (1) the patients enrolled had been diagnosed with lung cancer; (2) the results for the study included the correlation between Ki-67 and overall survival (OS) or disease-free survival (DFS); (3) the samples used in the studies were human lung tissue, serum or sputum but not animals or cell lines; (4) the techniques used to measure the expression level of Ki-67 in cancer tissue or tumors of the patients were immunohistochemistry (IHC), PCR/RT-PCR, ELISA or western blotting; (5) the study provided hazard ratios (HRs) and their 95% confidence intervals (CIs) or sufficient information for estimating these parameters; (6) the article was fully written in English or Chinese; and (7) the sample size was larger than 30. Studies were excluded if they met the following exclusion criteria:

### Table 4 Meta-regression for the OS and DFS analysis

| Variables       | HR     | Standard Error | t     | P > | Lower limit | Upper limit |
|-----------------|--------|----------------|-------|-----|-------------|-------------|
| OS              |        |                |       |     |             |             |
| Year            | 0.999  | 0.104          | −0.010| 0.991| 0.811       | 1.230       |
| Region          | 0.835  | 0.061          | −2.450| 0.017| 0.722       | 0.967       |
| Cancer type     | 0.954  | 0.045          | −0.990| 0.326| 0.868       | 1.049       |
| Sample size     | 1.108  | 0.122          | 0.930 | 0.353| 0.890       | 1.380       |
| Cutoff value    | 0.777  | 0.084          | −2.330| 0.022| 0.626       | 0.964       |
| Statistical method | 1.045  | 0.044          | 1.050 | 0.295| 0.961       | 1.137       |
| DFS             |        |                |       |     |             |             |
| Year            | 2.011  | 1.525          | 0.920 | 0.377| 0.379       | 10.678      |
| Region          | 0.591  | 0.289          | −1.070| 0.306| 0.201       | 1.736       |
| Cancer type     | 1.125  | 0.401          | 0.330 | 0.747| 0.513       | 2.467       |
| Sample size     | 0.793  | 0.348          | −0.530| 0.607| 0.302       | 2.081       |
| Cutoff value    | 0.767  | 0.363          | −0.560| 0.587| 0.271       | 2.172       |
| Statistical method | 0.994  | 0.221          | −0.030| 0.979| 0.609       | 1.622       |

### Table 5 Main results for meta-analysis between Ki-67 and clinicopathological features in lung cancer

| Clinicopathological features | Study(n) | Pooled OR(95%CI) | z     | P      | Heterogeneity | Publication bias |
|-----------------------------|----------|------------------|-------|--------|---------------|------------------|
| Age                         | 19       | 1.246(1.039–1.494) | 2.37  | 0.018  | 0.00%         | Fixed-effect     | 0.234 |
| Gender                      | 26       | 1.874(1.385–2.535) | 4.07  | < 0.001 | 69.70%        | Fixed-effect     | 1.000 |
| Histological type           | 16       | 0.397(0.236–0.667) | 3.49  | < 0.001 | 81.20%        | Fixed-effect     | 0.324 |
| Differentiation             | 11       | 1.993(1.262–3.146) | 2.96  | 0.003  | 66.30%        | Random-effect    | 0.893 |
| Pathologic stage            | 13       | 1.867(1.498–2.327) | 5.56  | < 0.001 | 23.10%        | Fixed-effect     | 1.000 |
| Tumor size                  | 12       | 1.436(1.127–1.29)  | 2.93  | 0.003  | 0.00%         | Fixed-effect     | 0.276 |
| Tumor stage                 | 11       | 1.287(0.882–1.877) | 1.31  | 0.191  | 55.30%        | Fixed-effect     | 0.086 |
| Metastasis                  | 4        | 2.609(0.667–10.204) | 1.38  | 0.168  | 64.50%        | Random-effect    | 0.428 |
| Lymph node                  | 23       | 1.653(1.285–2.127) | 3.91  | < 0.001 | 46.70%        | Random-effect    | 0.876 |
| TNM stage                   | 8        | 1.497(1.053–2.126) | 2.25  | 0.024  | 36.90%        | Fixed-effect     | 0.187 |
| Invasion                    | 3        | 0.993(0.511–1.930)  | 0.02  | 0.984  | 14.20%        | Fixed-effect     | 0.308 |
| Smoking                     | 15       | 3.087(2.504–3.806)  | 10.56 | < 0.001 | 39.40%        | Fixed-effect     | 0.711 |
criteria: (1) if they included animal experiments or cell lines or were pre-clinical studies, meta-analyses, reviews, comments, conference abstracts, letters or case reports; (2) articles in languages other than English or Chinese; and (3) studies did not include the key information for survival analyses such as HRs and 95% CIs. To avoid data duplication, when the same patient cohort was reported in different publications or the same article was found in different journals, only the most recent and complete publication was included.

Data extraction and quality assessment
All the articles were independently reviewed and selected by two investigators. Discrepancies were resolved by discussion and arbitrated by a third investigator. The following information was extracted from each publication: first author’s name, year of publication year, pathology type, tumor stage, number of patients, sample type, cut-off value of Ki-67, determination assay, method to extract HR and survival type. Additionally, we also obtained the clinicopathological characteristics of the lung cancer patients in the included studies including age (old/young), gender (male/female), histological type (adenocarcinoma/squamous carcinoma, ADC/SQC), smoking status (smoker/non-smoker), differentiation (poor/well or moderate), pathologic stage (III-IV/I-II), tumor size (large/small), tumor stage (T3–4/T1-T2), metastasis (yes/no), lymph node (N1-Nx/N0), TNM stage (T3–4/T1-T2) and invasion (yes/no). Especially, the tumor stage was used to describe the size and extent of tumor. And the TNM stage were used to define the progression of cancer based on the size and extension tumor, lymphatic involvement and metastasis status. Based on the Newcastle Ottawa Scale (NOS) criteria [142], studies with NOS scores higher than 6 are considered high-quality studies, whereas those with NOS scores less than 5 are defined as low-quality studies. This study was strictly performed according to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) [143] and the PRISMA checklist were also provided in the Additional file 8.

Statistical methods
HR and 95% CIs were used to measure the relationship between Ki-67 expression and prognosis of lung cancer patients. The most accurate determination was made when the study directly provided the HRs and 95% CIs. The multivariate HRs were calculated by using the Cox proportional hazards model, which could independently predict the survival outcome for the lung cancer patients. We preferentially chose the multivariate values when the study provided both univariate and multivariate HRs. If the above data were not available, we used the Engauge Digitizer version 4.1 to extract the survival rates from KM curves and estimated the HR according the method as Tierney et al. described [144, 145]. Moreover, we calculated the HR from the original survival data that the study provided using SPSS. An observed HR > 1 indicated worse prognosis for the lung cancer patients with high expression of Ki-67, if the 95% CIs for the overall HR was not 1, we considered the prognostic effect of Ki-67 on to be statistically significant. Odds ratio (OR) with 95% CIs were used to analyze the degree of association between Ki-67 level and clinicopathological characteristics. Heterogeneity was determined using the $\chi^2$ and inconsistency (I²) tests [146]. $I^2 > 50\%$ or $P < 0.1$ indicated substantial heterogeneity among the studies, in which case a random effects model was applied; otherwise, we utilized the fixed effects model. Then, subgroup analysis and meta-regression analysis were used to investigate any source of heterogeneity. Moreover, publication bias was assessed using Begg’s test and funnel plots, and $P$-values < 0.05 indicated statistically significant publication bias [147].
Additional files

**Additional file 1:** Figure S1. Forest plots for the relationships between Ki-67 expression and clinicopathological features of patients with lung cancer. A. Age B. Gender C. Histological type. (TIF 4444 kb)

**Additional file 2:** Figure S2. Forest plots for the relationships between Ki-67 expression and clinicopathological features of patients with lung cancer. A. Differentiation B. Pathologic stage C. Tumor size. (TIF 3759 kb)

**Additional file 3:** Figure S3. Forest plots for the relationships between Ki-67 expression and clinicopathological features of patients with lung cancer. A. Lymph node B. TNM stage C. Smoking. (TIF 3186 kb)

**Additional file 4:** Figure S4. Funnel plots for publication bias of clinicopathological features meta-analysis (A–C). A. Age B. Gender. C. Histological type. (TIF 20061 kb)

**Additional file 5:** Figure S5. Funnel plots for publication bias of clinicopathological features meta-analysis (A–C). A. Age B. Gender. C. Histological type. (TIF 21452 kb)

**Additional file 6:** Figure S6. Funnel plots for publication bias of clinicopathological features meta-analysis (G–J). A. Tumor stage B. Lymph node C. TNM stage. (TIF 21299 kb)

**Additional file 7:** Figure S7. Funnel plot for publication bias of clinicopathological features meta-analysis (Smoking status). (TIF 7465 kb)

**Additional file 8:** PRISMA 2009 Checklist. (DOC 63 kb)

**Abbreviations**

CDK1: Cyclin-dependent kinase1; CI: Confidence intervals; DFS: Disease-free survival; EGF: Epidermal growth factor receptor; HR: Hazard ratio; IHC: Immunohistochemistry; NSCLC: Non-small cell lung cancer; OR: Odds ratio; OS: Overall survival; SCLC: Small cell lung cancer; SQC: Squamous carcinoma

**Funding**

The current study was supported by the Funds of National Natural Science Foundation of China (NSFC 81360327, NSFC 81560469), Natural Science Foundation of Guangxi, China (2015GXNSFCA139009) and Guangxi Medical University Training Program for Distinguished Young Scholars (2017).

**Availability of data and materials**

The databases which we collection Literature include PubMed, Web of Science, EMBASE, and Chinese datasets (WanFang, China National Knowledge Infrastructure and Chinese VIP) until June 1, 2017.

**Authors’ contributions**

DMW and GC conceived of the project and designed the study. DMW and WJC wrote the manuscript. RMM and NZ participated in selecting study and extracting data. WJC and DYL take part in Manuscript Revise. GC was in charge of quality control. All authors read and approved the final manuscript.

**Ethics approval and consent to participate**

Not applicable.

**Consent for publication**

Not applicable.

**Competing interests**

The authors declare that they have no competing interests.

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