ON THE GEOMETRY OF MAXIMUM ENTROPY PROBLEMS

MICHELE PAVON † AND AUGUSTO FERRANTE ‡

Abstract. We show that a simple geometric result suffices to derive the form of the optimal solution in a large class of finite and infinite-dimensional maximum entropy problems concerning probability distributions, spectral densities and covariance matrices. These include Burg’s spectral estimation method and Dempster’s covariance completion, as well as various recent generalizations of the above. We then apply this orthogonality principle to the new problem of completing a block-circulant covariance matrix when an a priori estimate is available.

Key words. Maximum entropy problem, geometric principle, covariance selection, spectral estimation, Gibbs’ variational principle.

AMS subject classifications. 94A12, 90C46, 49K27, 60G10, 60J60, 62F30, 62H99

1. Prelude: Four famous maximum entropy problems. In this section, we briefly review four classical maximum entropy problems that have played an important role in the history of various scientific areas. These are namely problems where entropy is maximized under linear constraints. We shall later derive the form of the optimal solution in three of these problems by the same geometric principle (Theorem 3.3 in Section 3).

1.1. 1877: Boltzmann’s loaded dice. In 1877, Boltzmann [8, p.169] posed the following question: Consider \( N \) molecules that can only take the following \( p + 1 \) values of kinetic energy \( 0, \epsilon, 2\epsilon, \ldots, p\epsilon \). Suppose \( n_i \) molecules have kinetic energy \( i\epsilon, i = 0, 1, \ldots, p \). We then have a “macrostate”, a “Zustandverteilung” in Boltzmann’s language, indexed by \((n_0, n_1, \ldots, n_p)\) corresponding to the multinomial coefficient

\[
\frac{N!}{n_0!n_1!\ldots n_p!}
\]

“microstates” each having probability \((p+1)^{-N}\). Suppose that the sum of the kinetic energy of all molecules is a given quantity \(\lambda \epsilon = L\). Boltzmann proceeded to find the macrostate which corresponds to more microstates, namely that has highest probability, among those having total kinetic energy \(L\). This is, to the best of our knowledge, the first maximum entropy problem in history.

Boltzmann’s problem was popularized in the following form [65, 27]. Suppose \(N\) dice are rolled and we are informed that the total number of spots is \(N \cdot 4.5\). We are asked: What proportion of the dice are showing face \(i, i = 1, 2, \ldots, 6\)? The number of
different ways that $N$ dice can fall so that $n_i$ dice show face $i$ is given by

$$\frac{N!}{n_1!n_2!\ldots n_6!}, \quad \sum_{i=1}^{6} n_i = N. \quad (1.1)$$

Again, the “macrostate” $(n_1, n_2, \ldots, n_6)$ corresponds to $\frac{N!}{n_1!n_2!\ldots n_6!}$ “microstates” each having probability $6^{-N}$. To find the most probable macrostate, we need to maximize the multinomial coefficient (1.1) under the constraint

$$\sum_{i=1}^{6} i \cdot n_i = N \cdot 4.5. \quad (1.2)$$

This procedure will yield the macrostate, among those satisfying (1.2), that can be realized in more ways. Assuming that $N$ is large, we now use a crude version of Stirling’s approximation $N! \approx e^{-N} N^N$. We get

$$\frac{N!}{n_1!n_2!\ldots n_6!} \approx e^{-N} N^N \prod_{i=1}^{6} e^{-n_i \ln \left( \frac{n_i}{N} \right)} = e^{-\sum_{i=1}^{6} n_i \ln \left( \frac{n_i}{N} \right)} = e^{N H(p)}, \quad p_i = \frac{n_i}{N}, i = 1, 2, \ldots, 6.$$

Thus, for $N$ large, maximizing (1.1) under (1.2) is almost equivalent to maximizing the entropy

$$H(p) = -\sum_{i=1}^{6} p_i \ln (p_i) \quad (1.3)$$

under the constraint

$$\sum_{i=1}^{6} i \cdot p_i = 4.5.$$

The solution has the form

$$p_i^* = \frac{e^{\lambda_i}}{\sum_{i=1}^{6} e^{\lambda_i}}, \quad (1.4)$$

where the $\lambda_i$ must be such that

$$\sum_{i=1}^{6} i \cdot \frac{e^{\lambda_i}}{\sum_{i=1}^{6} e^{\lambda_i}} = 4.5.$$

Hence, the most probable macrostate is $(Np_1^*, Np_2^*, \ldots, Np_6^*)$ and we expect to find $n_i^* = Np_i^*$ dice showing face $i$. More is true: It can be shown [27, Chapter 13] that, for $N$ large, with probability close to one, other distributions satisfying (1.3) are close to $p^*$. This fact is sometimes referred to as Entropy Concentration Theorem [65]. More generally, when $F(p) := -\sum_k p_k \log(p_k)$, the maximizer of $F$ subject to a linear constraint $Lp = c$ has the form of a Boltzmann-Gibbs distribution

$$p_k = \frac{1}{Z} e^{-\langle \Lambda, L_k \rangle} \quad (1.5)$$

where $L_k$ is the $k$th column of the matrix $L$ and $Z$ a normalizing constant (partition function). This can of course also be formulated in the continuous setting (with integrals) and is also a basic result in statistics [31, 32, 33].
1.2. 1931: Schrödinger’s Bridges. In 1931/32, before the very foundations of probability were laid, Erwin Schrödinger studied the following abstract problem [83, 84]. Consider the evolution of a cloud of \( N \) independent Brownian particles. Here \( N \) is large, say of the order of Avogadro’s number. This cloud of particles has been observed having at some initial time \( t_0 \) an empirical distribution equal to \( \rho_0(x)dx \).

At some later time \( t_1 \), an empirical distribution equal to \( \rho_1(x)dx \) is observed which considerably differs from what it should be according to the law of large numbers, namely

\[
\left( \int_{t_0}^{t_1} p(t_0, y, t_1, x)\rho_0(y)dy \right) dx,
\]

where

\[
p(s, y, t, x) = \left[ 2\pi(t - s) \right]^{-\frac{n}{2}} \exp \left[ -\frac{|x - y|^2}{2(t - s)} \right], \quad s < t
\]

is the transition density of the Wiener process. It is apparent that the particles have been transported in an unlikely way. But of the many unlikely ways in which this could have happened, which one is the most likely? Schrödinger showed that the solution, namely the bridge from \( \rho_0 \) to \( \rho_1 \) over Brownian motion, has at each time a density \( q \) that factors as

\[
q(x, t) = \phi(x, t) \hat{\phi}(x, t),
\]

where \( \phi \) and \( \hat{\phi} \) solve the system

\[
\phi(t_0, x) = \rho_0(x) \quad \text{(1.6)}
\]

\[
\hat{\phi}(t_0, x) = \int p(t_0, y, t, x)\hat{\phi}(t_0, y)dy,
\]

\[
\phi(t_1, x) = \rho_1(x) \quad \text{(1.7)}
\]

\[
\hat{\phi}(t_1, x) = \int p(t_0, y, t, x)\phi(t_0, y)dy.
\]

It took more than fifty years before Föllmer, recovering Schrödinger’s original motivation, observed in [49] that this is a problem of large deviations of the empirical distribution on path space [43] connected, thanks to Sanov’s theorem [82], to a maximum entropy problem. Schrödinger’s problem may be considerably generalized. Let \( \Omega := C([t_0, t_1], \mathbb{R}^n) \) denote the family of \( n \)-dimensional continuous functions, let \( W_x \) denote Wiener measure on \( \Omega \) starting at \( x \), and let

\[
W := \int_{\mathbb{R}^n} W_x dx
\]

be stationary Wiener measure. Let \( \mathcal{D} \) be the family of distributions on \( \Omega \) that are equivalent to \( W \). For \( Q, P \in \mathcal{D} \), we define the relative entropy \( \mathbb{D}(P||Q) \) of \( P \) with respect to \( Q \) as

\[
\mathbb{D}(P||Q) = E_P[\log \frac{dP}{dQ}],
\]

where \( dP/dQ \) is the Radon-Nikodym derivative of \( P \) with respect to \( Q \). Let \( \mathcal{D}(\rho_0, \rho_1) \) be distributions in \( \mathcal{D} \) having the observed densities at times \( t_0 \) and \( t_1 \). If there is at least one \( P \) in \( \mathcal{D}(\rho_0, \rho_1) \) such that \( \mathbb{D}(P||Q) < \infty \), it may be shown that there exists

\( ^3 \)Large deviations theory has various applications in hypothesis testing, rate distortion theory, etc, see e.g. [27] Chapter 11, [38, 39] Chapters 2, 3, 7. For large deviations of the empirical distribution (level-2 large deviations) for diffusion processes see [15, 90] (see also [79] for a recent extension of this theory to discrete-time classical and quantum evolutions).
a unique minimizer $P_c$ in $\mathcal{D}(\rho_0, \rho_1)$ called in the language of Csiszár the I-projection of $Q$ onto $\mathcal{D}(\rho_0, \rho_1)$ [29, 30, 33]. It is the Schrödinger bridge from $\rho_0$ to $\rho_1$ over $Q$. In [33], using a conditional version of Sanov’s theorem established by Csiszár [30], it was shown that such I-projection $P_c$ provides the answer to Schrödinger’s original question: Namely, the asymptotic empirical distribution on path space, conditioned that the initial and final empirical distributions are $\rho_0(x)dx$ and $\rho_1(y)dy$, respectively, is indeed given by $P_c$.

1.3. 1967: Burg’s spectral estimation method. Suppose the covariance lags $c_k = \mathbb{E}[y(k)y(0)]$, $k = 0, 1, \ldots, n - 1$ of a stationary, zero-mean, Gaussian process have been estimated from the data. How should one extend the covariance? In 1967, while working on spectral estimation for geophysical data [11], Burg suggested the following approach. Rather than setting the other covariance lags to zero, one should set them to values such that they maximize the entropy rate (see Section 5 below) of the process. The solution is an autoregressive process of the form

$$y(m) = \sum_{k=1}^{n-1} a_k y(m-k) + w(m),$$

where $w$ is a zero-mean, Gaussian white noise sequence with variance $\sigma^2$. The parameters $a_1, \ldots, a_{n-1}, \sigma^2$ are such that the first $n$ covariance lags match the given ones.

1.4. 1972: Dempster’s covariance selection. In the seminal paper [37], a general strategy for completing a partially specified covariance matrix was introduced. Consider a zero-mean, multivariate Gaussian distribution with density

$$p(x) = (2\pi)^{-n/2} |\Sigma|^{-1/2} \exp \left\{ -\frac{1}{2} x^\top \Sigma^{-1} x \right\}, \quad x \in \mathbb{R}^n.$$ 

Suppose that the elements $\{\sigma_{ij}; 1 \leq i \leq j \leq n, (i, j) \in \bar{I} \}$, with $(i, i) \in \bar{I}$ for all $i = 1 \ldots n$, have been specified. How should $\Sigma$ be completed? Dempster resorts to a form of the Principle of Parsimony of parametric model fitting: As the elements $\sigma^{ij}$ of $\Sigma^{-1}$ appear as natural parameters of the model, one should set $\sigma^{ij}$ to zero for $1 \leq i \leq j \leq n, (i, j) \notin \bar{I}$. Notice that $\sigma^{ij} = 0$ has the probabilistic interpretation that the $i$-th and $j$-th components of the Gaussian random vector are conditionally independent given the other components [86]. We say that a positive definite completion $\Sigma^\circ$ of $\Sigma$ is a Dempster Completion if $[(\Sigma^\circ)^{-1}]_{i,j} = 0$ for all $(i, j) \notin \bar{I}$. In particular, Dempster proved that when a symmetric, positive-definite completion of $\Sigma$ exists, then there exists a unique Dempster’s Completion $\Sigma^\circ$. This completion maximizes the (differential) entropy

$$H(p) = -\int_{\mathbb{R}^n} \log(p(x))p(x)dx = \frac{1}{2} \log(\det \Sigma) + \frac{1}{2} n (1 + \log(2\pi))$$

among zero-mean Gaussian distributions having the prescribed elements $\{\sigma_{ij}; 1 \leq i \leq j \leq n, (i, j) \in \bar{I} \}$. Thus, Dempster’s Completion $\Sigma^\circ$ solves a maximum entropy problem, i.e. maximizes entropy under linear constraints.

2. Overture. The long tale of maximum entropy problems originates more than one hundred and thirty years ago with Boltzmann [8] at the dawn of statistical mechanics. Since then, several deep thinkers such as Jaynes [64, 65], Dempster [37],
Csiszár [31], to name but a few, have tried to explain the rationale behind the maximum entropy approach. Yet, this method, although never presented as a panacea [65, p.939], is still often viewed as an idiosyncratic choice. Before we get all tangled up with “predict states that can be realized by Nature in the greatest number of ways, while agreeing with your macroscopic information” (Jaynes interpreting Gibbs), apply the Principle of Parsimony of parametric model fitting (Dempster) or the axiomatic approach (Csiszár), we hasten reassure the reader: We are not going to give here even a précis of the motivation behind maximum entropy problems. Others have done it much better than we ever could. The scope of this paper is much more modest and yet, in a way, ambitious.

We want to point out that behind an endless string of maximum entropy solutions there is a simple geometric principle. Namely, that a whole class of seemingly unrelated results concerning probability distributions, spectral densities and covariance matrices are consequences of the same variational principle. All these problems feature linear constraints which determine an affine subspace $W$ in which the solution must be sought. Theorem 3.3 (or its generalization Theorem 9.1) simply states that the gradient (or a suitable generalization of it) of the entropy functional at a critical point must belong to the orthogonal complement (or, more generally, to the annihilator) of the subspace $V$ of which the affine space $W$ is a translation. Just to avoid any misunderstanding: We are not dealing here with the (usually challenging) existence problem [9, 10, 73, 74]. We simply want to derive in the most economic way the form of the optimal solutions assuming that they exist.

This orthogonality result is actually a direct consequence of a Lagrange multipliers argument. Nevertheless, we show that when the constraints are linear, there is no need to bring in our illustrious compatriot’s multipliers be they vectors, matrices or signals. One can simply skip the step, use this universal geometric result and presto! the form of the optimal solution appears. How can we have a geometric result when probability distributions/densities and spectra naturally belong to the intersection of suitable cones or simplices with $L^1$ spaces? The reader might look askance at this approach as, in general, in an infinite dimensional setting, $L^1$ spaces are not contained in $L^2$ spaces (one exception: absolutely summable sequences are also square summable). Hence, we simply don’t have the Euclidean or Hilbert space geometry where orthogonality makes sense. However, in many important maximum entropy problems, the solution together with an appropriate function of it (inverse, logarithm, etc.) also belongs to a suitable $L^2$ space (when this is not the case, see Section 9, a more general Banach space result may be applied). Thus, as we show, there is nothing to loose formulating the problem over an appropriate Hilbert space possibly intersected with a cone or a simplex.

One might wonder at this point: What has this to do with the well known orthogonality principle of linear quadratic optimization? Right on! Theorem 3.3 when applied to problems with quadratic criterion, yields well known results such as the orthogonality of the estimation error to the subspace generated by the available random variables in linear least-squares estimation. Thus, this orthogonality principle, a true deus ex machina, applies equally well to least-squares and entropic variational problems with linear constraints. Can this geometric result then be applied to any optimization problem in Hilbert space with linear constraints? Answer: No. The smoothness of the index functional is indispensable. For instance, the large and im-

---

4 This might well be the very reason that our simple observation has not been made before in a countless number of papers on maximum entropy problems.
important class of compressed sensing problems \cite{22, 39, 10, 19, 21, 20, 81} features as criteria $l^1$-type norms which do not even admit directional derivatives (they only admit one-sided directional derivatives as they are convex).

The reader might be doubtful by now: Don’t the authors of this paper know about information geometry, I-projections and the like \cite{25, 29, 88, 30, 1, 31, 2, 68, 5, 56, 91, 66, 78}? We do and are savvy enough to know that this body of work is of central importance in Mathematical Statistics, Information Theory, Signal Processing, Identification and Control. Our approach, however, is different. Rather than viewing the solution itself of maximum entropy problems as a projection in a suitable geometry and then developing a “Pythagorean Theorem for I-divergences”, our result involves usual orthogonality in Hilbert space (and the usual Pythagorean Theorem). Only that the orthogonality is a property of the differential of the entropy functional which does not in general relate to an “error”. In particular, our geometry does not depend on the particular entropic criterion employed but only on the Hilbert space in which the primal variables live.

The paper is outlined as follows. In Section 3, we present our basic variational result. This is then applied in Sections 4 and 5 to various classical and more recent maximum Burg’s entropy problems and in Section 6 to entropy problems with prior. In Section 7, we discuss maximum entropy problems on a finite measure space. In Section 8, we develop a new application to block-circulant covariance matrix completion when an a priori estimate is available. Finally, in Section 9, we give a generalization of our main result to Banach spaces.

3. Maxima on surfaces. Let $G : \mathbb{R}^3 \rightarrow \mathbb{R}$ be a continuously differentiable map and consider the surface (level set) $S \subset \mathbb{R}^3$ determined by the equation

$$G(x) = c, \quad c \in \mathbb{R}.$$  

Since the derivative of $G$ in the direction of a vector $v$ tangent to the surface $S$ must be zero, we get $\nabla G \cdot v = 0$. It namely follows the well-known fact that the gradient $\nabla G(x_0), x_0 \in S$, is perpendicular to the plane tangent to the surface $S$ at $x_0$. Let $F : \mathbb{R}^3 \rightarrow \mathbb{R}$ be another smooth functional and suppose that we are interested in maximizing $F$ over $S$. By the chain rule, at a local maximum point $x_0$, $\nabla F$ must be orthogonal to every differentiable curve on $S$ passing through $x_0$. We conclude that, at a maximum point $x_0$ the gradient of $F$ must also be perpendicular to the plane tangent to the surface $S$ at $x_0$ and therefore aligned with $\nabla G$, cf. e.g. \cite{42}, pp. 101-109). For instance, suppose we want to minimize $F(x, y, z) = x + y + z$ on the surface of the unit sphere $G(x, y, z) = x^2 + y^2 + z^2 = 1$. Since at a maximum point $\nabla F = (1, 1, 1)^T$ must be proportional to $\nabla G = 2(x, y, z)^T$, we conclude that maxima have equal components. It follows that the unique maximum point is $X_M = (3^{-1/2}, 3^{-1/2}, 3^{-1/2})$ ($X_m = -X_M$ is the unique minimum point).

The purpose of this paper is to show that a suitable generalization of this basic result is sufficient to derive the form of the optimal solution in a variety of maximum entropy problems.

In maximum entropy problems, the map $G$ is actually linear on a suitable vector space. Hence, $S = \text{ker } G + c$ is an affine space, namely the translation of the subspace $V = \text{ker } G$. In this case, the geometric principle simply says that, at a maximum point, $\nabla F$ must be perpendicular to the subspace $\text{ker } G$. We apply this geometric result to a large class of Burg-entropy and Shannon-entropy \cite{31, 32, 33} variational problems encompassing, as special cases, Burg spectral estimation method \cite{11, 12}, Dempster’s covariance completion \cite{37} and Gibbs-like variational principles \cite{13}. In
Burg’s maximum entropy problems, one maximizes the (integral of) the logarithm of a positive quantity, be it a probability density, a spectrum or the determinant of a positive definite matrix, under linear constraints. The latter determine the affine space \(W\). Theorem 3.3 simply says that the Fréchet differential of the entropy functional at a critical point must belong to the orthogonal complement of the subspace \(V\) of which the affine space \(W\) is a translation (coset). In the Burg’s entropy case, this entails that the adjoint of the inverse of the solution must belong to \(V^\perp\). In the case of Shannon maximum entropy problems, the orthogonality condition concerns the logarithm of the solution.

Classical results can then be readily re-derived and generalized. For example, our result contains the key for the (considerable) recent generalizations developed in \[52, 44, 46, 23, 47, 48\]. The case when a prior estimate is available is also covered by this geometric principle. In the Burg’s case the entropic functional turns into a multivariate Itakura-Saito divergence \[4, 60\]. In the Shannon case, entropy is replaced by the Kullback-Leibler divergence (relative entropy) \[71\]). As an application, we show how our result can be used to extend the results of \[23\] to the case when a prior estimate of the circulant covariance is available. The latter problem deals with identifying of the parameters of a stationary reciprocal process given the first covariance lags and an a priori covariance estimate.

Let \(H\) be a Hilbert space and let \(F : H \to \mathbb{R}\) be a functional. We say that \(F\) is Gâteaux-differentiable at \(h_0\) in direction \(v\) if the limit

\[
F'(h_0; v) := \lim_{\epsilon \to 0} \frac{F(h_0 + \epsilon v) - F(h_0)}{\epsilon}
\]

exists. In this case, \(F'(h_0; v)\) is called the directional derivative of \(F\) at \(h_0\) in direction \(v\). We say that \(F\) is Fréchet-differentiable at \(h_0\) if there exists an element \(DF(h_0)\) in \(H\) such that

\[
\lim_{\|h\|_H \to 0} \frac{|F(h_0 + h) - F(h_0) - \langle DF(h_0), h \rangle_H|}{\|h\|_H} = 0.
\]

The element \(DF(h_0)\) is called the Fréchet differential of \(F\) at \(h_0\). Fréchet differentiability is stronger than Gâteaux differentiability. In fact, we have the following result \[70\] p.50).

**Proposition 3.1.** Let \(F\) be Fréchet differentiable at \(h_0\). Then, \(DF(h_0)\) is unique and, for any \(v \in H\), \(F\) is Gâteaux differentiable at \(h_0\) in direction \(v\) and it holds

\[
F'(h_0; v) = \langle DF(h_0), v \rangle_H.
\]

Conversely, when \(F\) is Gâteaux differentiable on an open set \(U \subseteq H\) and its Gâteaux derivative is linear and continuous at each point of \(U\) then \(F\) is Fréchet differentiable in \(U\). Finally, when \(F\) is convex, if it is Gâteaux differentiable in all directions \(v\) then it is Fréchet differentiable.

In some applications, we cannot expect that the functional be Fréchet differentiable at the point of interest. We may, however, have that a formula like \(3.1\) holds when \(v\) varies over a subspace. More precisely, let \(V \subseteq H\) be a (not necessarily closed) subspace and \(h \in H\). Consider the corresponding coset \(W := h + V\) which is an affine space over \(V\). Observe that, for \(w \in W\) and \(v \in V\), \((w + \epsilon v) \in W\) for all for all real \(\epsilon\), namely \(w\) is an internal point of \(W\) in direction \(v\).

**Definition 3.2.** We say that \(w_c\) is a critical point of \(F\) over \(W = h + V\) if \(F'(w_c; v) = 0\) for all \(v \in V\).
Theorem 3.3. Let \( \mathcal{W} := h + \mathcal{V} \) be an affine space. Assume that the functional \( F \) is Gâteaux-differentiable at \( w_c \in \mathcal{W} \) in any direction \( v \in \mathcal{V} \) and that the Gâteaux differential is given by the linear, continuous map \( F'(w_c; v) = (D_v F(w_c), v)_{\mathcal{H}} \) where \( D_v F(w_c) \in \mathcal{H} \). Then \( w_c \) is a critical point of \( F \) over \( \mathcal{W} \) if and only if \( D_v F(w_c) \in \mathcal{V}^\perp \). When \( F \) is actually Fréchet differentiable at \( w_c \in \mathcal{W} \), \( w_c \) is critical if and only if \( DF(w_c) \in \mathcal{V}^\perp \).

Proof. \( F'(w_c; v) = 0 \) for all \( v \in \mathcal{V} \) if and only if \( (D_v F(w_c), v)_{\mathcal{H}} = 0, \forall v \in \mathcal{V} \). \( \square \)

4. Matricial variational problems.

4.1. Geometric result. Let \( \mathcal{H} = \mathbb{C}^{n \times n} \) (or \( \mathcal{H} = \mathbb{R}^{n \times n} \)) be the space of \( n \times n \) matrices endowed with the inner product \( \langle M_1, M_2 \rangle := \text{tr}[M_1^* M_2] \), where \( * \) denotes transposition plus conjugation (we write \( M^{-*} \) for \( (M^{-1})^* \)). The following result was established in [46].

Lemma 4.1. Let

\[
F(M) := \log |\det [M]|.
\] If \( M \) is nonsingular then, for all \( \delta M \in \mathcal{H} \)

\[
F'(M; \delta M) = \text{tr} [M^{-1} \delta M] = (M^{-*}, \delta M).
\]

It now follows from Proposition [3.1] that \( F \) is Fréchet differentiable in the open set of non-singular matrices and

\[
DF(M) = M^{-*}.
\]

We are interested in extremizing \((1.1)\) over an affine space, namely a coset of the form \( \mathcal{W} = A + \mathcal{V} \), where \( A \in \mathcal{H} \) and \( \mathcal{V} \) is a subspace of \( \mathcal{H} \).

Theorem 4.2. Let \( \mathcal{W} = A + \mathcal{V} \) be an affine space. Then a nonsingular matrix \( M_c \in \mathcal{W} \) extremizes \( F(M) = \log |\det [M]| \) over \( \mathcal{W} \) if and only if \( M_c^{-*} \in \mathcal{V}^\perp \).

Proof. Let \( M_c \in \mathcal{W} \) be non-singular. By \((4.3)\), we have \( DF(M_c) = M_c^{-*} \). The conclusion now follows from Theorem 3.3. \( \square \)

4.2. Dempster’s covariance selection. In various applications, index \((4.1)\) must be extremized (or rather maximized) on the intersection between an affine space \( \mathcal{W} \) and a convex cone. A typical example is that of the cone of positive semidefinite matrices. This is the case considered by Dempster in the seminal paper [37] where a general strategy for completing a partially specified covariance matrix was introduced.

We now show that Theorem 4.2 provides a geometrical interpretation of one of the key features of Dempster’s result. To see this, consider the Dempster’s problem with the same notation as in Subsection 1.4. Let \( \mathcal{W} \) be the affine space of symmetric matrices having elements \( \{\sigma_{ij}; 1 \leq i \leq j \leq n, (i, j) \in \mathcal{I}\} \). Notice that \( \mathcal{W} \) is affine over the subspace \( \mathcal{V} \) of symmetric matrices having zeros in the positions \( \mathcal{I} \). Observe next that the solution \( \Sigma \) is constrained to be in the intersection between \( \mathcal{W} \) and the convex cone of positive definite matrices. On this set, maximizing the index \((4.1)\) or the entropy \((1.8)\) is equivalent. Thus, the two criteria yield the same solution. Moreover, \((i, i) \in \mathcal{I} \) for all \( i = 1 \ldots n \), i.e. \( [\Sigma]_{ii} \) are all fixed so that \( [\Sigma]_{ij} \leq \sqrt{[\Sigma]_{ii}[\Sigma]_{jj}} \) and hence the feasible set is bounded. Finally, as \( \Sigma \) tends to be singular, i.e. it approaches the boundary of the cone, \( H(p) \) tends to \(-\infty \) which implies that the solution can be searched among positive definite matrices. Thus, under the feasibility assumption, the optimal solution exists and lies in the interior of the cone. We can then repeat locally
the argument of Theorem 4.2 to conclude that the maximum entropy completion $\Sigma_c$ is such that $\Sigma_c^{-1} \in \mathcal{V}^\perp$. Finally, observe that $\mathcal{V}^\perp$ is the space of matrices having zeros in $\mathcal{I}$, the complement of $\mathcal{I}$. Indeed, let $e_i$ denote the $i$-th canonical vector in $\mathbb{R}^n$ and observe that for $(i, j) \in \mathcal{I}$, the rank one matrix $e_i e_j^\top$ belongs to $\mathcal{V}$. If $M \in \mathcal{V}^\perp$, we must have

$$0 = \text{tr} \left[ (e_i e_j^\top) M \right] = \text{tr} [e_j e_i^\top M] = e_j^\top M e_i = [M]_{ij}, \quad \forall (i, j) \in \mathcal{I}.$$ 

Thus, the maximum entropy completion $\Sigma_c$ is a Dempster’s completion.

### 4.3. General matrix completion

In [46], Dempster’s completions were shown to solve suitable entropy-like variational problems for general nonsingular matrices. Again, the form of the extremal completions (no uniqueness is there guaranteed) when they exist is provided by Theorem 4.2.

### 5. Matricial functions

#### 5.1. The orthogonality result

Consider now the Hilbert space $\mathcal{H}$ of square integrable functions taking values in the space of $m \times m$ Hermitian matrices. We denote by $\mathbb{H}_n$ the $n^2$-dimensional, real vector space of Hermitian matrices of dimension $n \times n$. Hence, $\mathcal{H} = L^2(\mathcal{T}, \mathbb{H}_m)$ with scalar product

$$\langle \Phi, \Psi \rangle_\mathcal{H} := \frac{1}{2\pi} \int_{-\pi}^\pi \text{tr} \left[ \Phi(e^{j\theta}) \Psi(e^{j\theta}) \right] d\theta.$$

Consider the functional

$$(5.1) \quad F(\Phi) = \frac{1}{2\pi} \int_{-\pi}^\pi \log |\text{det} [\Phi(e^{j\theta})]| d\theta.$$

**Lemma 5.1.** Suppose $\Phi \in L^\infty(\mathcal{T}, \mathbb{H}_n)$ is coercive. Then, for any $\delta \Phi \in L^\infty(\mathcal{T}, \mathbb{H}_n)$ the directional derivative of $F$ exists and is given by the linear map

$$(5.2) \quad F'(\Phi; \delta \Phi) = \frac{1}{2\pi} \int_{-\pi}^\pi \text{tr} \left[ \Phi^{-1}(e^{j\theta}) \delta \Phi(e^{j\theta}) \right] d\theta = \langle \Phi^{-1}, \delta\Phi \rangle_\mathcal{H}.$$

**Proof.** Observe that, for $\delta \Phi \in L^\infty(\mathcal{T}, \mathbb{H}_n)$ and $|\varepsilon|$ sufficiently small, $\Phi(e^{j\theta}) + \varepsilon \delta \Phi(e^{j\theta})$ is a.e. positive definite. After bringing the derivative under the integral sign, we can use Lemma 4.1 for almost all $\theta$. \hfill \Box

Let $\mathcal{W} = A + \mathcal{V}$ be an affine space in $L^\infty(\mathcal{T}, \mathbb{H}_n)$, namely $A \in L^\infty(\mathcal{T}, \mathbb{H}_n)$ and $\mathcal{V}$ is a subspace of $L^\infty(\mathcal{T}, \mathbb{H}_n)$.

Then Theorem 5.2 yields:

**Theorem 5.2.** Let $\mathcal{W} = A + \mathcal{V}$ be as above and $\Phi \in L^\infty(\mathcal{T}, \mathbb{H}_n)$ coercive. Then, if $\Phi_c$ is a critical point of $\Phi$ over $\mathcal{W}$, we have $\Phi_c^{-1} \in \mathcal{V}^\perp$.

**Proof.** By Lemma 5.1 if such a $\Phi_c$ extremizes (5.1), then $\langle \Phi_c^{-1}, v \rangle_\mathcal{H} = 0$ for all $v \in \mathcal{V}$. Namely, $\Phi_c^{-1} \in \mathcal{V}^\perp$. \hfill \Box

---

5Actually, the case of full-rank rectangular matrices, with the Moore-Penrose pseudoinverse in the place of the inverse, was also treated in [46].

6$\Phi$ is called coercive if $\exists \alpha > 0$ s.t. $\Phi(e^{j\theta}) - \alpha I_m$ is a.e. positive definite on $\mathcal{T}$.

7For a not necessarily closed subspace $\mathcal{V}$ of $L^2(\mathcal{T}, \mathbb{H}_n)$, the orthogonal complement $\mathcal{V}^\perp$ is the closed subspace of $u \in L^2(\mathcal{T}, \mathbb{H}_n)$ such that $\langle u, v \rangle_{L^2} = 0, \forall v \in \mathcal{V}$. 
5.2. Burg’s maximum entropy covariance extension. In his seminal work [11, 12], Burg introduced a spectral estimation method based on the maximization of entropy which is widely used in signal processing. We now show that Theorem 5.2 provides a most transparent reason why the solution has to be an AR process. Consider a discrete-time Gaussian process \( \{y_k; k \in \mathbb{Z}\} \) taking values in \( \mathbb{R}^m \). Let \( Y_{[-n,n]} \) be the random vector obtained by considering the window \( y_{-n}, y_{-n+1}, \ldots, y_0, \ldots, y_{n-1}, y_n \), and let \( p_{Y_{[-n,n]}} \) denote the corresponding joint density. The differential entropy rate of \( y \) is defined by

\[
h_r(y) := \lim_{n \to \infty} \frac{1}{2n + 1} H(p_{Y_{[-n,n]}}, \quad \text{(5.3)}
\]

if the limit exists, where \( H(p_{Y_{[-n,n]}}) \) denotes the entropy of the density of the random vector \( Y_{[-n,n]} \), cf. (1.8). In [69], Kolmogorov established the following important result.

**Theorem 5.3.** Let \( y = \{y_k; k \in \mathbb{Z}\} \) be a \( \mathbb{R}^m \)-valued, zero-mean, Gaussian, stationary, purely nondeterministic of full rank process with spectral density \( \Phi_y \). Then

\[
h_r(y) = \frac{m}{2} \log(2\pi e) + \frac{1}{4\pi} \int_{-\pi}^{\pi} \log \det \Phi_y(e^{i\vartheta}) d\vartheta. \quad \text{(5.4)}
\]

As is well-known, there is also a fundamental connection between the quantity appearing in (5.4) and the optimal one-step-ahead predictor: The multivariate Szegő-Kolmogorov formula reads

\[
\det R = \exp \left\{ \frac{1}{2\pi} \int_{-\pi}^{\pi} \log \det \Phi_y(e^{i\vartheta}) d\vartheta \right\}, \quad \text{(5.5)}
\]

where \( R \) is the error covariance matrix corresponding to the optimal predictor. Consider now the multivariate covariance extension problem. Let \( C_k, k = 0, 1, \ldots, n-1 \) of dimension \( m \times m \) be some estimated covariance lags of an unknown stationary process \( y \). Then Burg’s problem consists in finding a stationary process \( y \) with spectral density \( \Phi_y \) which maximizes the index

\[
F(\Phi_y) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \log \det \Phi_y(e^{i\vartheta}) d\vartheta. \quad \text{(5.6)}
\]

among all spectral densities having as first \( n \) Fourier coefficients \( C_k, k = 0, 1, \ldots, n-1 \). In view of Kolmogorov’s result (5.4), maximizing the entropy rate of a stationary Gaussian process is equivalent to maximizing the integral of \( \log \det \Phi_y \). Assume that the block-Toeplitz matrix \( \Sigma_n \)

\[
(5.7) \quad \Sigma_n = \begin{bmatrix}
C_0 & C_1 & \cdots & C_{n-1} \\
C_1^* & C_0 & \cdots & C_{n-2} \\
\vdots & \vdots & \ddots & \vdots \\
C_{n-1}^* & C_{n-2}^* & \cdots & C_0
\end{bmatrix}
\]

is positive definite. Then [61] there are infinitely many spectra having the prescribed Fourier coefficients.

\[\text{Actually, the solution to this problem maximizes the entropy rate in the larger class of second-order processes [20].}\]
Consider now the matrix pseudo-polynomial \( P(e^{j\vartheta}) = \sum_{k=-n+1}^{n-1} C_k e^{-j\vartheta k} \), with \( C_{-k} := C_k^* \), and define the subspace \( \mathcal{V}_n \) of \( L^\infty(T,\mathbb{H}_n) \) of functions whose Fourier coefficients \( R_i \) vanish for all \( i = -n+1, \ldots, n-1 \) and obey to the symmetry constraint \( R_i = R_{-i}^* \). Then the constraint in Burg’s problem can be expressed as \( \Phi \in \mathcal{W} \cap \mathcal{S} \), where the affine space \( \mathcal{W} \) is defined by
\[
\mathcal{W} = P + \mathcal{V}_n
\]
and \( \mathcal{S} \) is the convex cone of bounded, coercive spectral densities. On \( \mathcal{S} \), (5.1) and (5.6) coincide, and \( F \) is strictly concave. Thus, an extremizer \( \Phi_c \) is actually a maximum point. By Theorem 5.2, this maximum point \( \Phi_c \) is such that \( \Phi_{-1}^c \in \mathcal{V}_n^\perp \). Observe now that \( \mathcal{V}_n^\perp \) is given by the matricial polynomials of the form
\[
Q(e^{j\vartheta}) = \sum_{k=-n+1}^{n-1} A_k e^{-j\vartheta k}, \quad A_{-k} = A_k^*.
\]
We conclude that the optimal spectrum has the form
\[
\Phi_c(e^{j\vartheta}) = \left[ \sum_{k=-n+1}^{n-1} A_k e^{-j\vartheta k} \right]^{-1}, \quad A_{-k} = (A_k^*)^*
\]
for some matrices \( A_k, k = -n + 1, \ldots, 0, \ldots, n - 1 \) which permit to satisfy the constraints on the first \( n \) coefficients. Thus, the solution process is an AR process. If only some of the \( C_k, k = 0, 1, \ldots, n - 1 \) are available, the classical approach to the problem requires a certain effort and some ad hoc reasoning to get the solution form. Theorem 5.2, on the contrary, yields immediately that in (5.8) \( A_k^* = 0 \) for all \( k \) corresponding to missing \( C_k \)'s.

5.3. A more general moment problem. We consider next a generalization of Burg’s problem studied by Byrnes, Georgiou and Lindquist and co-workers [16, 14, 17, 50, 53, 57, 51, 76, 59] in the frame of generalized moment problems. In their broad research effort, having applications, besides spectral estimation, to robust control problems, elements of a parametric family of rational spectral densities were recognized from the start [16, 15] to be critical points of logarithmic entropy-like functionals.

Consider a transfer function
\[
G(z) = (zI - A)^{-1}B, \quad A \in \mathbb{C}^{n \times n}, B \in \mathbb{C}^{n \times m}, \quad n > m,
\]
where \( A \) has all its eigenvalues in the open unit disk, \( B \) has full column rank, and \( (A, B) \) is a reachable pair.\(^9\) Suppose \( G(z) \) models a bank of filters fed by a wide sense stationary, purely nondeterministic, \( \mathbb{C}^m \)-valued process \( y \):

\[\begin{array}{ccc}
     \text{y(t)} & \text{G(z)} & \text{x(t)} \\
\end{array}\]

\(^9\)A pair \( (A, B) \) is called reachable in Systems Theory [67] if the matrix \( [B \mid AB \mid \ldots \mid A^{n-1}B] \) has full row rank.
Let $x$ be the $n$-dimensional stationary output process
\begin{equation}
  x_{k+1} = Ax_k + By_k, \quad k \in \mathbb{Z}.
\end{equation}
We denote by $\Sigma$ the covariance of $x_k$. The spectrum $\Phi$ must then satisfy the following moment constraint
\begin{equation}
  \frac{1}{2\pi} \int_{-\pi}^{\pi} G(e^{i\vartheta}) \Phi(e^{i\vartheta}) G^*(e^{i\vartheta}) d\vartheta = \Sigma.
\end{equation}
As in [17, 57, 51, 45, 48], we now consider the problem of determining spectral densities $\Phi$ satisfying (5.11) for a given $\Sigma > 0$. The covariance extension is a special case of this problem corresponding to $G(z) := [z^{-n} I \mid z^{-n+1} I \mid \ldots \mid z^{-1} I]$ and $\Sigma$ equal to the Toeplitz matrix in (5.7). More details on this fact may be found in [57] where other classical problems are shown to be special cases of the above. The most important of these problems is the celebrated Nevanlinna-Pick interpolation problem of fundamental importance in various $H^\infty$ control problems [41, 18, 7, 58].

We now show how to treat this problem in our geometric framework. Let, as before, $\mathcal{H} = L^2(\mathbb{T}, \mathbb{H}_m)$. Consider now the linear operator
\begin{equation}
  \Gamma : L^\infty(\mathbb{T}, \mathbb{H}_m) \to \mathbb{H}_n,
  \Phi \mapsto \frac{1}{2\pi} \int_{-\pi}^{\pi} G(e^{i\vartheta}) \Phi(e^{i\vartheta}) G^*(e^{i\vartheta}) d\vartheta.
\end{equation}
It follows that for the constraint (5.11) to be feasible, $\Sigma$ must belong to the linear space
\begin{equation}
  \text{Range } \Gamma := \left\{ M \in \mathbb{H}_n \mid \exists \Phi \in L^\infty(\mathbb{T}, \mathbb{H}_m) \text{ such that } \frac{1}{2\pi} \int_{-\pi}^{\pi} G(e^{i\vartheta}) \Phi(e^{i\vartheta}) G^*(e^{i\vartheta}) d\vartheta = M \right\}.
\end{equation}
Consider now the following generalization of Burg’s problem: Maximize the entropy index (5.6) subject to (5.11) where $\Sigma$ is assumed to be positive definite. Suppose that (5.11) is feasible, namely there exists a spectral density $\Phi_0 \in L^\infty(\mathbb{T}, \mathbb{H}_m)$ satisfying this constraint. Then, the family $\mathcal{W}$ of hermitian-valued functions satisfying (5.11) may be expressed as
\begin{equation}
  \mathcal{W} = \Phi_0 + \mathcal{V},
\end{equation}
where $\mathcal{V} = \{ \Phi \in L^\infty(\mathbb{T}, \mathbb{H}_m) \mid G \Phi G^* = 0 \}$. In other words, $\mathcal{V} = \ker \Gamma$. The constraint in the generalized Burg problem can be expressed as $\Phi \in \mathcal{W} \cap \mathcal{S}$, where $\mathcal{S}$ is the convex cone of bounded, coercive spectral densities. Since
\begin{equation}
  \langle \int_{-\pi}^{\pi} G\Phi G^* d\vartheta, M \rangle_{\mathbb{H}_n} := \text{tr} \left[ \int_{-\pi}^{\pi} G\Phi G^* d\vartheta \right] M = \text{tr} \left[ \int_{-\pi}^{\pi} \Phi G^* MG d\vartheta \right] = \langle \Phi, G^* MG \rangle_{\mathcal{H}},
\end{equation}
we have that the adjoint of $\Gamma$, mapping $\mathbb{H}_n$ to $L^\infty(\mathbb{T}, \mathbb{H}_m)$, is given by
\begin{equation}
  \Gamma^* : M \mapsto G^* MG.
\end{equation}
In particular, $\text{Range } \Gamma^* = \{ \Phi = G^* MG, M \in \mathbb{H}_n \} \subset C(\mathbb{T}, \mathbb{H}_m)$ the continuous Hermitian-valued functions on the unit circle. Since $\text{Range } \Gamma^*$ is finite-dimensional, it is necessarily closed and we have
\begin{equation}
  \mathcal{V}^\perp = [\ker \Gamma]^\perp = \text{Range } \Gamma^* = \{ \Phi = G^* MG, M \in \mathbb{H}_n \}.
\end{equation}
By Theorem 5.2, the maximum point $\Phi_c$ is such that $\Phi_c^{-1} \in V^\perp$. Hence, the optimal spectrum has the form

$$\Phi_c(e^{j\vartheta}) = \left[ G(e^{j\vartheta})^* \Lambda_c G(e^{j\vartheta}) \right]^{-1},$$

for some Hermitian $\Lambda_c$ such that $G(e^{j\vartheta})^* \Lambda_c G(e^{j\vartheta}) > 0$ on $T$ and the constraint (5.11) is satisfied, namely

$$\int_{-\pi}^{\pi} G [G^* \Lambda_c G]^{-1} G^* \frac{d\vartheta}{2\pi} = \Sigma.$$

Indeed, Georgiou showed in [52] that the unique solution of the generalized Burg problem has the form (5.16) with

$$\Lambda_c = \Sigma^{-1} B (B^* \Sigma^{-1} B)^{-1} B^* \Sigma^{-1}. \quad (5.17)$$

### 6. Variational entropy problems with “prior”.

#### 6.1. Matricial problems.

Consider now the same set up as in Section 4, where a “prior” nonsingular estimate $N$ of the matrix $M$ is available. Rather than extremizing (maximizing) (4.1), we now consider the problem of finding a matrix belonging to the given affine set $W$ and which extremizes the index

$$F(M) := \log |\det [N]| - \log |\det [M]| + \text{tr} \left( N^{-1} M \right) \quad (6.1)$$

(see below for insights and motivation for this choice). Lemma 4.1 now becomes:

**Lemma 6.1.** Let $F(M)$ be given by (6.1). If $M$ is nonsingular then for any $\delta M \in H = \mathbb{C}^{n \times n}$,

$$F'(M; \delta M) = \text{tr} \left[ (-M^{-1} + N^{-1}) \delta M \right], \quad (6.2)$$

and $DF(M) = -M^{-1} + N^{-1}$. By Theorem 3.3, we get:

**Theorem 6.2.** Let $W = A + V$ be an affine set in $H = \mathbb{C}^{n \times n}$. Let $N$ be a nonsingular matrix in $H$. Then the nonsingular matrix $M_c \in W$ extremizes (6.1) over $W$ if and only if $(M_c^{-*} - N^{-*}) \in V^\perp$.

In order to motivate the choice (6.1), we first recall a few basic facts on entropy for Gaussian random vectors and processes that may be found e.g. in [80, 63, 27]. The *relative entropy* or *Kullback-Leibler* pseudo-distance or *divergence* between two probability densities $p$ and $q$, with the support of $p$ contained in the support of $q$, is defined by

$$D(p \parallel q):= \int_{\mathbb{R}^n} p(x) \log \frac{p(x)}{q(x)} \, dx, \quad (6.3)$$

see e.g [27]. In the case of two zero-mean Gaussian densities $p$ and $q$ with positive definite covariance matrices $M$ and $N$, respectively, the relative entropy is given by

$$D(p \parallel q) = \frac{1}{2} \left[ \log \det (M^{-1} N) + \text{tr} (N^{-1} M) - n \right]. \quad (6.4)$$

Hence, when $N$ and $M$ are positive definite, minimizing index (6.1) is indeed equivalent to minimizing the Kullback-Leibler divergence between two Gaussian random
vectors which is one of the central problems in statistical modeling. Indeed, as is well-known, (6.4), originates from maximum likelihood considerations, cf. e.g. [13] Section II. An important application of this result is the estimation of a structured covariance matrix. In this class of problems we need to estimate a covariance matrix $\Sigma$ in such a way that it satisfies some linear constraints. The sample covariance estimate $\hat{\Sigma}$ will normally fail to satisfy the given linear constraints so that the problem of computing $\hat{\Sigma}_c$ that satisfies the constraints and is as close as possible to $\Sigma$, arises naturally, see [13, 53, 47, 78] for more details and applications. In particular, in [47] the constraint is given by $\Sigma \in \text{Range } \Gamma$ (as defined in (5.13)). It was shown there (Proposition 3.2) that

$$\mathcal{V} = \{ \Sigma : (I - \Pi_B)(\Sigma - A\Sigma A^*)(I - \Pi_B) = 0 \},$$

with $\Pi_B$ being the orthogonal projection onto $\text{im } (B)$, so that it is easy to see that

$$\mathcal{V}^\perp = \{ \Delta = (I - \Pi_B)\Lambda(I - \Pi_B) - A^*(I - \Pi_B)\Lambda(I - \Pi_B)A : \Lambda \in \mathbb{H}_n \}.$$

Then, Theorem 6.2 can be used to get in a straightforward manner the form of the optimal $\hat{\Sigma}_c$ presented in [47] Section IV:

$$\hat{\Sigma}_c = \left( \Sigma^{-1} + (I - \Pi_B)\Lambda(I - \Pi_B) - A^*(I - \Pi_B)\Lambda(I - \Pi_B)A \right)^{-1}, \Lambda \in \mathbb{H}_n.$$

6.2. Matricial functions problems with “prior”. As much as Theorem 4.2 also Theorem 6.2 may be generalized to the case when $\mathcal{H} = L^2(\mathbb{T}, \mathbb{H}_n)$. In this setting, we consider $\Phi \in L^\infty(\mathbb{T}, \mathbb{H}_n)$ coercive and a given “prior” $\Psi$ also essentially bounded and coercive. The index to be extremized is

$$F(\Phi, \Psi) = \frac{1}{2\pi} \int_{-\pi}^{\pi} \left\{ \log(\det \Psi) - \log(\det \Phi) + \text{tr } [\Psi^{-1}\Phi] \right\} d\vartheta.$$

Motivation for considering this index will be provided after the statement of the next result. A straightforward generalization of Lemma 6.1 and Theorem 3.3 now give a result germane to Theorem 6.2.

**Theorem 6.3.** Let $\mathcal{H}$ be as before $L^2(\mathbb{T}, \mathbb{H}_n)$ and let $\mathcal{V} = \mathcal{A} + \mathcal{V}$ be an affine set in $L^\infty(\mathbb{T}, \mathbb{H}_n)$ and $\Phi_1(\cdot, \vartheta)^\perp \subset \mathcal{W}$ be coercive. Then $\Phi_1$ extremizes (6.8) over $\mathcal{W}$ if and only if $(\Phi_1^{-1} - \Psi)^{-1} \in \mathcal{V}^\perp$.

To provide some motivation and insight for index (6.8), we consider two zero-mean, jointly Gaussian, stationary, purely nondeterministic processes $y = \{y_k; k \in \mathbb{Z} \}$ and $z = \{z_k; k \in \mathbb{Z} \}$ taking values in $\mathbb{R}^m$. We consider the relative entropy rate $\mathbb{D}_r(y\|z)$ between $y$ and $z$ defined as

$$\mathbb{D}_r(y\|z) := \lim_{n \to \infty} \frac{1}{2n + 1} \mathbb{D}(p_{Y_{-n:n}}\|p_{Z_{-n:n}})$$

where $p_{Y_{-n:n}}$ and $p_{Z_{-n:n}}$ are the densities of the random vectors obtained from $y$ and $z$, respectively, by considering the “windows” from time $-n$ to time $n$. Following in his mentor’s footsteps, the great information theorist M. Pinsker [80] proved the following important result (see also [87], [63], [77]):

**Theorem 6.4.** Let $y = \{y_k; k \in \mathbb{Z} \}$ and $z = \{z_k; k \in \mathbb{Z} \}$ be $\mathbb{R}^m$-valued, zero-mean, Gaussian, stationary, purely nondeterministic processes with spectral density functions $\Phi_y$ and $\Phi_z$, respectively. Assume, moreover, that at least one of the following conditions is satisfied:
1. $\Phi_y \Phi_z^{-1}$ is bounded;
2. $\Phi_y \in L^2(-\pi, \pi)$ and $\Phi_z$ is coercive.

Then

$$D_r(y \parallel z) = \frac{1}{4\pi} \int_{-\pi}^{\pi} \left\{ \log \det (\Phi_y^{-1}(e^{j\vartheta}) \Phi_z(e^{j\vartheta})) + \text{tr} \left[ \Phi_z^{-1}(e^{j\vartheta}) (\Phi_y(e^{j\vartheta}) - \Phi_z(e^{j\vartheta})) \right] \right\} d\vartheta.$$  \hfill (6.10)

The index (6.10) has the form of a multivariate Itakura-Saito divergence of speech processing and is basically the same as (6.8). Indeed, one of the main results of [48] is based on the minimization of (6.8) where $\Psi$ is a given “prior” spectral density and $\Phi$ must belong to the intersection between the cone $S$ of positive definite spectral densities and the affine set $W$ of the solutions of the moment problem (5.11), for given $G$ and $\Sigma$. Since the constraint is as before, so are the spaces $W$ and $V$. In particular, we have

$$V^\perp = \{ \Phi = G^*MG, M \in \mathbb{H}_n \}.$$  

By Theorem 6.3, we get the form of the optimal spectrum derived in [48]

$$\Phi_c = (\Psi^{-1} + G^*\Lambda_c G)^{-1}, \quad \Lambda_c \in \mathbb{H}_n,$$  

where $\Lambda_c$ permits to satisfy (5.11).

6.3. Kullback-Leibler approximation of spectral densities. Consider the same set up as in Subsection 5.3 in the scalar case ($m = 1$) when an a priori estimate of the spectrum $\Psi$ is available. The latter is assumed to be essentially bounded and coercive. In [57], the following constrained approximation problem was studied: Minimize $F(\Phi) = D(\Psi \parallel \Phi) = \int \log (\Psi/\Phi) \Psi$ among coercive spectra $\Phi \in L^\infty(T)$ satisfying (5.11). Notice that minimization occurs with respect to the second argument. This permits to include the maximum entropy in this framework ($\Psi \equiv 1$) and to obtain a rational solution rather than in the exponential class when $\Psi$ is rational. Further justification for this choice of the criterion may be found in [57]. In this case, for $\delta \Phi \in L^\infty$, $F'(\Phi; \delta \varphi) = -\langle \Phi^{-1}(\Psi, \delta \varphi) \rangle_{L^2}$. Since the constraint is as in (5.11), so is the space $V^\perp$, see (5.15). By Theorem 3.3 we conclude that the optimal spectrum has the form obtained in [57]

$$\Phi_c(e^{j\vartheta}) = \frac{\Psi(e^{j\vartheta})}{G^*(e^{j\vartheta})\Lambda_c G(e^{j\vartheta})}, \quad \Lambda_c \in \mathbb{H}_n.$$  

The difficulties of extending this result to the multivariable case are illustrated in [55, p.1062].

7. Shannon entropy for finite measure spaces. The Shannon entropy underlying all the criteria so far considered will be here addressed directly via the first (rather than the second) part of equation (1.8) and with a finite measure $\mu$ replacing Lebesgue measure. Let $(X, \mathcal{X}, \mu)$ be a finite measure space and let $\varphi_i, i = 1, \ldots, d$ be functions in $H = L^2(X, \mathcal{X}, \mu)$ and $\alpha \in \mathbb{R}^d$. Consider the problem of finding a nonnegative function $p$ in $L^\infty(X, \mathcal{X}, \mu)$ maximizing the Shannon entropy

$$F(p) = H(p) = -\int_X \log[p(x)]p(x)d\mu$$  \hfill (7.1)
under the constraints

\[(7.2) \quad \int_X p(x) d\mu = 1,\]
\[(7.3) \quad \int_X \varphi_i(x) p(x) d\mu = \alpha_i, \quad i = 1, \ldots, d.\]

Lemma 5.1 can be readily adapted to this setting. Let \(p_c \in L^\infty(X, \mathcal{X}, \mu)\) be nonnegative and bounded away from zero \(\mu\) a.e. Let \(\delta p \in L^\infty(X, \mathcal{X}, \mu)\). Then the directional derivative of the functional (7.1) in direction \(\delta p\) exists at \(p_c\) and is given by

\[F'(p_c; \delta p) = \int_X [-1 + \log p_c(x)] \delta p(x) d\mu = \langle -1 + \log p_c, \delta p \rangle_H.\]

Let us show that the fundamental geometric result Theorem 3.3 provides the form of the extremal solution also in this case. Suppose there exists \(p_0 \in L^\infty(X, \mathcal{X}, \mu)\) a.e. everywhere positive satisfying (7.2)-(7.3). Then \(p \in L^\infty(X, \mathcal{X}, \mu)\) also satisfies the constraints if it belongs to the affine space \(p_0 + \mathcal{V}\) where \(\mathcal{V}\) is the subspace of functions \(f \in L^\infty(X, \mathcal{X}, \mu)\) such that

\[(7.4) \quad \int_X f(x) d\mu = 0,\]
\[(7.5) \quad \int_X \varphi_i(x) f(x) d\mu = 0, \quad i = 1, \ldots, d.\]

Observe now that \(\mathcal{V}^\perp\) is the subspace of functions of the form \(\vartheta_0 + \sum_{i=1}^d \vartheta_i \varphi_i(x)\). Observe also that for \(p_c\) bounded and bounded away from zero as above, \(\log p_c\) also belongs to \(L^\infty(X, \mathcal{X}, \mu)\) and, consequently, to \(L^2(X, \mathcal{X}, \mu)\). By Theorem 3.3 we conclude that \((-1 + \log p_c) \in \mathcal{V}^\perp\), it must namely be of the form

\[(7.6) \quad p_c(x) = C \exp \left[ \sum_{i=1}^d \vartheta_i \varphi_i(x) \right]\]

for some values \(C\) and \(\vartheta_i, i = 1, \ldots, d\) that permit to satisfy the constraints. This is just the well-known fact that, if the maximizer exists, it belongs to the exponential family. In the case when \(d = 1\) and \(\varphi_1 = H\) the Hamiltonian function, we get a baby version of Gibbs variational principle, namely that the Gibbs distribution

\[p_G(x) = C \exp \left[ -\frac{H(x)}{kT} \right]\]

minimizes the free energy \(\langle H, p \rangle - kTF(p)\) where \(F\) is as in (7.1), \(k\) is Boltzmann’s constant and \(T\) is absolute temperature. [43].

The well know fact that among all probability densities with given mean and variance the Gaussian has maximum entropy can also be derived in this framework by taking as “reference” measure \(\mu\) a Gaussian measure.

8. Reciprocal processes identification with prior. In this section, we consider the problem of block-circulant covariance completion addressed in [23, 24] and we show that our result allows for a direct solution of this more general problem also in the case (not considered there) when a prior estimate is available. The above mentioned block-circulant covariance completion is equivalent to the computation of
the parameters of a stationary reciprocal process of order \( n \) defined on the discrete circle \( \mathbb{Z}/N\mathbb{Z} \). A process \( y(t) \) defined on \( \mathbb{Z}/N\mathbb{Z} \) is reciprocal if it enjoys the following property. Take any two points \( i, j \in \mathbb{Z}/N\mathbb{Z} \): They divide the discrete circle into two (discrete) arcs. Then process \( y(t) \) is reciprocal of order 1 if \( y(t) \) and \( y(\tau) \) are conditionally independent given \( y(i) \) and \( y(j) \), for any \( i, j \) and for any \( t \) and \( \tau \) belonging to different arcs. The process \( y(t) \) is reciprocal of order \( n \) if \( y(t) \) and \( y(\tau) \) are conditionally independent given \( y(i), y(i+1), \ldots y(i+n-1) \) and \( y(j), y(j+1), \ldots y(j+n-1) \), for any \( i, j \) and for any \( t \) and \( \tau \) belonging to different arcs. Reciprocal processes defined on (a finite interval of) the integer line can be seen as a special class of discrete Markov random fields restricted to one dimension. Stationary reciprocal processes defined on \( \mathbb{Z}/N\mathbb{Z} \) are potentially useful for describing signals which naturally live in a finite region of the time (or space) line such as texture images.

Let \( \Sigma_i \in \mathbb{R}^{m \times m}, i = 0, 1, \ldots, n \) be given. In [23] the problem has been considered to compute the parameters of a stationary reciprocal process of order \( n \) defined on the discrete circle \( \mathbb{Z}/N\mathbb{Z} \) such that the first \( n+1 \) covariance lags of this process match the given \( \Sigma_i \), \( i = 0, 1, \ldots, n \). For the importance and applications of this problem we refer to [23] and references therein. For a discussion of stationary reciprocal processes, we refer to [75]. In [23] is was shown that this problem is equivalent to compute an extension \( \Sigma_i \in \mathbb{R}^{m \times m}, i = n+1, n+2, \ldots, N-1 \) in such a way that the symmetric block-Toeplitz matrix \( \Sigma \) whose first block row is \( \left[ \Sigma_0 \mid \Sigma_1^T \mid \ldots \Sigma_{N-1}^T \right] \) maximizes

\[
F(\Sigma) := \log[\det(\Sigma)]
\]

in the set \( \mathcal{W} \cap \mathcal{S} \), where \( \mathcal{S} \) is the cone of positive definite matrices and \( \mathcal{W} \) is the affine space of block-circulant symmetric matrices such that the north-west corner block of dimension \( m(n+1) \times m(n+1) \) is equal to the symmetric block-Toeplitz matrix \( \Sigma_{11} \) whose first block row is \( \left[ \Sigma_0 \mid \Sigma_1^T \mid \ldots \Sigma_{n}^T \right] \). The form of solution to this problem may be easily computed by using Theorem 4.2. In fact, define

\[
U = \begin{bmatrix}
0 & I_m & 0 & \ldots & 0 \\
0 & 0 & I_m & \ldots & 0 \\
\vdots & \vdots & \ddots & \vdots & \vdots \\
0 & 0 & 0 & \ldots & I_m \\
I_m & 0 & 0 & \ldots & 0
\end{bmatrix} \in \mathbb{R}^{Nn \times Nm},
E = \begin{bmatrix}
I_m & 0 & \ldots & 0 \\
0 & I_m & \ldots & 0 \\
0 & 0 & \ddots & \vdots \\
\vdots & \vdots & \ddots & 0 \\
0 & 0 & \ldots & I_m \\
\end{bmatrix} \in \mathbb{R}^{Nm \times (n+1)m}.
\]

where \( I_m \) denotes the \( m \times m \) identity matrix. Clearly, \( U^T U = U U^T = I_{mN} \); i.e. \( U \) is orthogonal. Note that a matrix \( C \) with \( N \times N \) blocks is block-circulant if and only if it commutes with \( U \), namely if and only if it satisfies

\[
U^T C U = C.
\]

The affine set \( \mathcal{W} \) may be then characterized as

\[
\mathcal{W} = \{ \Sigma = \Sigma^T : E^T \Sigma E = \Sigma_{11}, U^T \Sigma U = \Sigma \} = A + \mathcal{V}
\]

with \( A \in \mathcal{W} \) and

\[
\mathcal{V} := \{ \Sigma = \Sigma^T : E^T \Sigma E = 0, U^T \Sigma U = \Sigma \}.
\]

It is not difficult to check that

\[
\mathcal{V} \perp = \{ \Delta = E \Lambda E^T + U \Theta U^T - \Theta, \Lambda = \Lambda^T \in \mathbb{R}^{(n+1)m \times (n+1)m}, \Theta = \Theta^T \in \mathbb{R}^{Nm \times Nm} \}.
\]
Hence the optimal solution, if it exists, has the form

\[ \Sigma_c = (E\Lambda E^T + U\Theta U^T - \Theta)^{-1}, \]

where \( \Lambda = \Lambda^T \in \mathbb{R}^{(n+1)m \times (n+1)m} \), and \( \Theta = \Theta^T \in \mathbb{R}^{Nm \times Nm} \) must be chosen in such a way that the constraints are satisfied. This can be done through convex duality as discussed in [23]. The dual problem consists here in the unconstrained maximization of the concave function

\[ L(\Lambda, \Theta) = \text{tr} \log (E\Lambda E^T + U\Theta U^T - \Theta) + \text{tr} I - \text{tr} (\Lambda \Sigma^{-1}), \]

over a suitable set of multiplier pairs \((\Lambda, \Theta)\). Once the optimal parameters \( \Lambda \) and \( \Theta \) have been found, the optimal solution has inverse \( \Sigma_c^{-1} \) which is a block-circulant matrix whose first block-row has the form

\[ [M_0 | M_1 | \ldots | M_n | 0 | 0 | \ldots | 0 | M_n^T | M_{n-1}^T | \ldots | M_1^T], \]

where the matrices \( M_i \) are the sought for parameters of the stationary reciprocal process.

We now address the case when a prior information is available in terms of the parameters of a reciprocal process (possibly of higher order), or, equivalently of a prior positive definite covariance matrix \( \Sigma_p \in \mathbb{R}^{Nm \times Nm} \). In this case, instead of maximizing (8.1) we minimize the divergence (see (6.4))

\[ F(\Sigma) := \left[ \log \det (\Sigma^{-1} \Sigma_p) + \text{tr} (\Sigma_p^{-1} \Sigma) \right] \]

under the same constraints. By employing Theorem 6.2 we get the form of the optimal solution is

\[ \Sigma_c = (E\Lambda E^T + U\Theta U^T - \Theta + \Sigma_p^{-1})^{-1}, \]

where, again, \( \Lambda = \Lambda^T \in \mathbb{R}^{(n+1)m \times (n+1)m} \), and \( \Theta = \Theta^T \in \mathbb{R}^{Nm \times Nm} \) must be chosen in such a way that the constraints are satisfied. As before, this can be done by solving a dual problem for which existence can be proven along the lines of [23]. From (8.9) it follows that when \( \Sigma_p \) is also the covariance matrix of a stationary reciprocal process of order \( n \) or less, the optimal solution is also reciprocal of order \( n \) and coincides with the optimal solution of the problem without prior! This remarkable result follows from (8.9) and the fact that there exists a unique block circulant covariance completion satisfying the linear constraints and having block zeros in the first row as in (8.7). If instead, \( \Sigma_p \) is the covariance matrix of a stationary reciprocal process of order \( n_1 > n \) (requiring a larger memory), then the optimal solution is the covariance of a reciprocal process of order \( n_1 \) whose parameters may be read in the first block-row of \( \Sigma_c^{-1} \).

9. Extension to functionals defined on a Banach space. In some applications, Theorem 3.3 does not suffice. For this reason, we mention the straightforward extension of our main result to functionals \( F \) defined on a Banach space. Let \( X \) be a Banach space and let \( F : X \to \mathbb{R} \) be a functional. We say that \( F \) is \( Gâteaux\)-differentiable at \( x_0 \) in direction \( v \) if the limit

\[ F'(x_0; v) := \lim_{\epsilon \to 0} \frac{F(x_0 + \epsilon v) - F(x_0)}{\epsilon} \]
exists. In this case, \( F'(x_0; v) \) is called the directional derivative of \( F \) at \( x_0 \) in direction \( v \). We say that \( F \) is Fréchet-differentiable at \( x_0 \) if there exists a bounded linear functional on \( \mathcal{X} \) \( DF_{x_0} \) such that

\[
\lim_{\|x\|_{\mathcal{X}} \to 0} \frac{|F(x_0 + x) - F(x_0) - DF_{x_0}(h)|}{\|x\|_{\mathcal{X}}} = 0.
\]

The functional \( DF_{x_0} \) is called the Fréchet differential of \( F \) at \( x_0 \). Again, if \( F \) is Fréchet differentiable at \( x_0 \), then \( DF_{x_0} \) is unique and, for any \( x \in \mathcal{X} \), \( F \) is Gâteaux differentiable at \( x_0 \) in direction \( v \) and it holds

\[
(9.1) \quad F'(x_0; v) = DF_{x_0}(v).
\]

**Theorem 9.1.** Let \( \mathcal{X} \) be a Banach space, let \( \mathcal{V} \subseteq \mathcal{X} \) be a subspace, let \( x \in \mathcal{X} \) and consider the corresponding coset \( \mathcal{W} := x + \mathcal{V} \). Assume that the functional \( F \) is Fréchet-differentiable at \( w_c \in \mathcal{W} \). Then \( w_c \) is a critical point of \( F \) over \( \mathcal{W} \) if and only if \( DF_{w_c} \) belongs to the annihilator of \( \mathcal{V} \).

**Proof.** Observe that \( F'(w_c; v) = 0 \) for all \( v \in \mathcal{V} \) if and only if \( DF_{w_c}(v) = 0, \forall v \in \mathcal{V} \).

\[ \square \]

When \( F \) is not Fréchet-differentiable at \( w_c \) but merely Gâteaux differentiable in directions varying in a subspace, a generalization such as in Theorem 3.3 can be established.

**10. Closing comments.** In this paper, we have established a simple orthogonality condition that allows to derive the form of the optimal solution in a plethora of maximum entropy problems. We feel that this geometric condition affords a considerable conceptual simplification allowing to cast least-squares and maximum entropy problems in the same framework (admittedly, not as deep as the one provided in [31]). It can, moreover, be readily generalized to abstract situations and to problems with nonlinear constraints. Further study is needed to see whether this approach may be suitably adapted to the abstract setting of Subsection 1.2. A suitable mixture of the geometry we have seen in Burg’s and in Dempster’s problems in Subsections 5.2 and 4.2 might provide the key to understanding AR and ARMA Identification of Graphical Models, a topic which has recently received considerable attention, see e.g. [72, 34, 32, 85, 3]. Finally, we should never forget the motto over the entrance to Plato’s Academy: “Αγεωρέτριτος μηδείς εισήγητω”, namely “Let no one untrained in geometry enter.”

**Acknowledgments.** The authors wish to thank two anonymous reviewers for a careful reading and for providing several constructive suggestions. In particular, we are thankful to one reviewer for suggesting to employ a generalization without Fréchet differentiability of the main result and for encouraging us to access Boltzmann’s original work [3].

**REFERENCES**

[1] S. Amari, *Differential-geometrical methods in statistics*, Lecture notes in statistics, Springer-Verlag, Berlin, 1985
[2] S. Amari and H. Nagaoka, *Methods of information geometry*, Translations of Mathematical Monographs; v. 191, American Mathematical Society, 2000.
[3] E. Avventi, *Spectral Moment Problems: Generalizations, Implementation and Tuning*, PhD thesis, KTH, Stockholm, Sweden, 2011.
[4] M. Basseville. Distance Measures for Signal Processing and Pattern Recognition. *Signal Processing*, 18:349–369, 1989.

[5] R. Bhatia, Positive definite matrices. Princeton Univ Press, 2007.

[6] P. Billingsley, *Convergence of Probability Measures*, Wiley, New York, 1968.

[7] A. Blomqvist, A. Lindquist, and R. Nagamune. Matrix-valued Nevanlinna-Pick interpolation with complexity constraint: An optimization approach. *IEEE Trans. Aut. Control*, 48:2172–2190, 2003.

[8] L. Boltzmann, Über die Beziehung zwischen dem zweiten Hauptsatze der mechanischen Wärmetheorie und der Wahrscheinlichkeitsrechnung resp. den Sätzen über das Wärmegleichgewicht. Wiener Berichte 76, 373-435, 1877. Reprinted in F. Hasenöhrl (ed.): Wissenschaftliche Abhandlungen. Leipzig: J. A. Barth 1909, Vol. 2, 164-223.

[9] J. Borwein and A. Lewis, Duality relationships for entropy-like minimization problems, *SIAM J. Control Optim.*, 29 (1991) 325-338.

[10] J. Borwein and A. Lewis, Partially-finite programming in L1 and the existence of maximum entropy estimates, *SIAM J. Optim.*, 3 (1993) 248-267.

[11] J. P. Burg, Maximum entropy spectral analysis, in *Proc.37th Meet. Society of Exploration Geophysicists*, 1967. Reprinted in Modern Spectrum Analysis, D. G. Childers, Ed. New York: IEEE Press, 1978. pp. 34-41.

[12] J. P. Burg, Maximum entropy spectral analysis, Ph.D.dissertation, Dept. of Geophysics, Stanford University, Stanford, CA,1975.

[13] J. Burg, D. Luenberger, and D. Wenger, Estimation of Structured Covariance Matrices, *Proceedings of the IEEE*, 70, 963–974, 1982.

[14] C. I. Byrnes, T. Georgiou, and A. Lindquist, A new approach to spectral estimation: A tunable high-resolution spectral estimator, *IEEE Trans. Sig. Proc.*, 49, 3189–3205, 2000.

[15] C. I. Byrnes, T. Georgiou, and A. Lindquist, A generalized entropy criterion for Nevanlinna-Pick interpolation with degree constraint, *IEEE Trans. Aut. Control*, 46, 822–839, 2001.

[16] C. I. Byrnes, S. Gusev, and A. Lindquist, A convex optimization approach to the rational covariance extension problem, *SIAM J. Control and Optimization*, 37, 211–229, 1999.

[17] C. I. Byrnes, S. Gusev, and A. Lindquist. From finite covariance windows to modeling filters: A convex optimization approach. *SIAM Review*, 43:645–675, 2001.

[18] C. I. Byrnes, T. Georgiou, and A. Lindquist. A generalized entropy criterion for Nevanlinna-Pick interpolation with degree constraint: A convex optimization approach to certain problems in systems and control. *IEEE Trans. Aut. Control*, 46:822–839, 2001.

[19] E. J. Candès and Y. Plan, Matrix completion with noise, *Proceedings of the IEEE*, Vol. 98(6): 925 – 936, 2010.

[20] E. J. Candès and B. Recht, Exact matrix completion via convex optimization, *Found. of Comput. Math.*, 9 717–772, 2009.

[21] E. J. Candès and J. Romberg, Sparsity and incoherence in compressive sampling, *Inverse Problems*, 23 (3) pp. 969-985, 2007.

[22] E. J. Candès, J. Romberg and T. Tao, Robust uncertainty principles: Exact signal reconstruction from highly incomplete frequency measurement, *IEEE Trans. on Information Theory*, 52 (2) pp. 489 – 509, February 2006.

[23] F. Carli, A. Ferrante, M. Pavon, and G. Picci, A Maximum Entropy solution of the Covariance Extension Problem for Reciprocal Processes, *IEEE Trans. Aut. Control*, 56, Issue 9, September 2011, 1999-2012.

[24] F. Carli and T. Georgiou, On the Covariance Completion Problem under a Circulant Structure, *IEEE Trans. Aut. Control*, 56 (4), April 2011, 918 - 922.

[25] N. N. Chentsov, *Statistical decision rules and optimal inference* (in Russian), Nauka, 1972. Translations of Mathematical Monographs, Amer. Math. Soc., 1982, no. 53.

[26] B. S. Choi and T. M. Cover, An Information-Theoretic Proof of Burg’s Maximum Entropy Spectrum, *Proc. of the IEEE*, VOL. 72 (8) (1984), 1094-1095.

[27] T. M. Cover and J. A. Thomas. *Elements of Information Theory*, Wiley, New York, 1991.

[28] H. Cramér, H., Sur un nouveau théorème-limite de la théorie des probabilités, *Actualités Sci. Indust.* 736 (1938), 5-23.

[29] I. Csiszár, I-divergence geometry of probability distributions and minimization problems, *Annals of Probability*, 3, pp. 146-158, 1975.

[30] I. Csiszár, Sanov property, generalized 1-projections, and a conditional limit theorem, *Annals of Probability*, 12, pp. 768-793, 1984.

[31] I. Csiszár, “Why least squares and maximum entropy? An axiomatic approach to inference for linear inverse problems.” *The Annals of Statistics*, 19(4): 2032-2006, 1991.

[32] I. Csiszár and F. Matúš (2008), On minimization of entropy functionals under moment constraints. Proceedings ISIT 2008, Toronto, Canada, 2101-2105.
[33] I. Csiszár and F. Matúš. Information projections revisited. IEEE Trans. Inform. Theory, 49:1474-1490, 2003.
[34] R. Dahlhaus, Graphical interaction models for multivariate time series, Metrika, 51, pp. 157-172, 2000.
[35] D. Dawson, L. Gorostiza and A. Wakolbinger, Schrödinger processes and large deviations, J. Math. Physics, 31 (10), 2385-2388, 1990.
[36] A. Dembo and O. Zeitouni, Large deviations techniques and applications, Jones and Bartlett Publishers, Boston, 1993.
[37] A. P. Dempster, Covariance selection, Biometrics, 28,157–175, 1972.
[38] J. D. Deuschel and D. Stroock, Large Deviations, Academic Press, Boston, 1989.
[39] D. Donoho, Compressed sensing, IEEE Trans. on Information Theory, 52 (4), pp.1289 - 1306, April 2006.
[40] D. Donoho, For most large underdetermined systems of linear equations, the minimal ell-1 norm solution is also the sparsest solution, Communications on Pure and Applied Mathematics, 59(6), pp. 797-829, June 2006.
[41] J. Doyle, B. Francis, and A. Tannenbaum. Feedback Control Theory. Macmillan Publishing Company, 1992.
[42] C. H. Edwards, Advanced Calculus of Several Variables, Academic Press, New York, 1973.
[43] R. S. Ellis, Entropy, Large deviations and statistical mechanics, Springer-Verlag, New York, 1985.
[44] P. Enqvist and J. Karlsson. Minimal itakura-saito distance and covariance interpolation. In 47th IEEE Conference on Decision and Control, CDC 2008., pages 137 –142, 9-11 2008.
[45] A. Ferrante, M. Pavon, and F. Ramponi. Hellinger vs. Kullback-Leibler multivariable spectrum approximation. IEEE Trans. Aut. Control, 53:954–967, 2008.
[46] A. Ferrante and M. Pavon, Matrix Completion à la Dempster by the Principle of Parsimony, IEEE Trans. Information Theory, Vol. 57:3925–3931, June 2011.
[47] A. Ferrante, M. Pavon, and M. Zorzi. A maximum entropy enhancement for a family of high-resolution spectral estimators. IEEE Trans. Aut. Control, 57, Issue 2, 318–329, 2012.
[48] A. Ferrante, C. Masiero and M. Pavon, Time and spectral domain relative entropy: A new approach to multivariate spectral estimation, March 2011, arXiv:math-ph/1103.5602v2, IEEE Trans. Aut. Control, to appear.
[49] H. Föllmer, Random fields and diffusion processes, in: École d’Été de Probabilités de Saint-Flour XV-XVII, edited by P. L. Hennequin, Lecture Notes in Mathematics, Springer-Verlag, New York, 1988, vol.1362,102-203.
[50] T. Georgiou, Spectral estimation by selective harmonic amplification, IEEE Trans. Aut. Control 46, 29–42, 2001.
[51] T. Georgiou, The structure of state covariances and its relation to the power spectrum of the input, IEEE Trans. Aut. Control, 47:1056–1066, 2002.
[52] T. Georgiou. Spectral analysis based on the state covariance: the maximum entropy spectrum and linear fractional parameterization. IEEE Trans. Aut. Control, 47:1811–1823, 2002.
[53] T. Georgiou, Spectral analysis based on the state covariance: the maximum entropy spectrum and linear fractional parameterization, IEEE Trans. Aut. Control, 47, 1811–1823, 2002.
[54] T. Georgiou, “Structured covariances and related approximation questions,” in Directions in Mathematical Systems Theory and Optimization (A. Rantzer and C. Byrnes, eds.), vol. 286 of Lecture Notes in Control and Information Sciences, pp. 135–140, Springer Berlin / Heidelberg, 2003.
[55] T. Georgiou, Relatıve entropy and the multivariable multidimensional moment problem. IEEE Trans. Inform. Theory, 52:1052–1066, 2006.
[56] T. Georgiou, Distance and Riemannian metrics for spectral density functions, IEEE Transactions on Signal Processing , vol. 55 (8), pp. 3995-4003, 2007.
[57] T. Georgiou and A. Lindquist. Kullback-Leibler approximation of spectral density functions. IEEE Trans. Inform. Theory, 49:2910–2917, 2003.
[58] T. Georgiou and A. Lindquist. Remarks on control design with degree constraint. IEEE Trans. Aut. Control, AC-51:1150–1156, 2006.
[59] T. Georgiou and A. Lindquist, A convex optimization approach to ARMA modeling, IEEE Trans. Aut. Control, AC-53, 1108–1119, 2008.
[60] R. Gray, A. Buzo, A. Jr Gray, and Y. Matsuyama. Distortion measures for speech processing, IEEE Trans. Acoustics, Speech and Signal Proc., 28:367–376, 1980.
[61] U. Grenander and G. Szegö, Toeplitz Forms and Their Applications, University of California Press, Berkeley, CA, 1958.
[62] S. Hojsgaard and S. L. Lauritzen, Graphical Gaussian models with edge and vertex symmetries. J. of Royal Statistical Society, Series B, 70, 1005-1027, 2008.
[63] S. Ihara. *Information Theory for Continuous Systems*. World Scientific, Singapore, 1993.

[64] E. T. Jaynes, *Information Theory and Statistical Mechanics*, Physical Review Series II, 106 (4): 620630, 1957. doi:10.1103/PhysRev.106.620. MR87305, and *Information Theory and Statistical Mechanics II*, Physical Review Series II, 108 (2): 171190, 1957. doi:10.1103/PhysRev.108.171. MR96414.

[65] E. T. Jaynes. On the rationale of maximum-entropy methods. Proceedings of the IEEE, 70(9):939–952, Sept. 1982.

[66] X. Jiang, L. Ning, and T. Georgiou. Distances and riemannian metrics for multivariate spectral densities, June 2011. *IEEE Trans. Aut. Contr.*. to appear.

[67] E. T. Jaynes. *On the rationale of maximum-entropy methods*. Proceedings of the IEEE, 70(9):939–952, Sept. 1982.

[68] X. Jiang, L. Ning, and T. Georgiou. Distances and riemannian metrics for multivariate spectral densities, June 2011. *IEEE Trans. Aut. Contr.*. to appear.

[69] R. E. Kalman, P. L. Falb and M.A. Arbib, *Topics in Mathematical System Theory*, McGraw-Hill, New York, 1969.

[70] R. Kass and P. Vos, *Geometrical Foundations of Asymptotic Inference*, Wiley, New York, 1997.

[71] A. N. Kolmogorov. On the Shannon theory of information in the case of continuous signals. *IRE Trans. Inform. Theory*, 2:102–108, 1956.

[72] S. L. Lauritzen, *Graphical Models*, Oxford University Press, 1996.

[73] C. Leonard, Minimizers of energy functionals under not very integrable constraints, *J. Convex Anal*. 10 (2003) 63-68.

[74] C. Leonard, Minimization of entropy functionals, *J. Math. Anal. Appl*. 346 (2008) 183-204.

[75] B. Levy and A. Ferrante. Characterization of Stationary Discrete-Time Gaussian Reciprocal Processes over a Finite Interval. *SIAM J. Matrix Analysis*. Vol. 24(2):334–355, 2002.

[76] A. Lindquist, *Prediction-error approximation by convex optimization*, in *Modeling, Estimation and Control: Festschrift in honor of Giorgio Picci on the occasion of his sixty-fifth Birthday*, A. Chiuso, A. Ferrante and S. Pinzoni (eds), Springer-Verlag, pp. 265-275, 2007.

[77] A. Lindquist and G. Picci. *Linear Stochastic Systems: A Geometric Approach to Modeling, Estimation and Identification*. In preparation: preprint available in http://www.math.kth.se/~alq/LPbook.

[78] L. Ning, X. Jiang, and T. Georgiou. Geometric methods for estimation of structured covariances. Preprint, Oct. 2011, available at arXiv:1110.3695v1 2011.

[79] M. Pavon and F. Ticozzi, Discrete-time classical and quantum Markovian evolutions: Maximum entropy problems on path space, *J. Math. Phys.*., 51, 042104-042125 (2010).

[80] M. S. Pinsker. *Information and information stability of random variables and processes*. Holden-Day, San Francisco, 1964. Translated by A. Feinstein.

[81] J. Romberg, Imaging via compressive sampling, *IEEE Signal Processing Magazine*, 25 (2), pp. 14 - 20, March 2008.

[82] I. S. Sanov, On the probability of large deviations of random magnitudes (in Russian), Mat. Sb. N. S., 42 (84) (1957) 1144. Select. Transl. Math. Statist. Probab., 1, 213-244 (1961).

[83] E. Schrödinger, Über die Umkehrung der Naturgesetze, Sitzungsberichte der Preuss Akad. Wissen. Berlin, Phys. Math. Klasse (1931), 144-153.

[84] E. Schrödinger, Sur la théorie relativiste de l’´electron et l’interpretation de la m´ecanique quantique, Ann. Inst. H. Poincaré 2, 269 (1932).

[85] J. Songsiri, J. Dahl, L. Vandenberghe, Graphical models of autoregressive processes. In: Y. Eldar and D. Palomar, editors, *Convex Optimization in Signal Processing and Communications*, Cambridge University Press (2010), 89-116.

[86] T. P. Speed and H. T. Kiiveri, Gaussian Markov distributions over finite graphs, *The Annals of Statistics*, 14 (1986), 138-150.

[87] A. A. Stoorvogel and J. H. Van Schuppen. System identification with information theoretic criteria. In S. Bittanti and G. Picci, editors, *Identification, Adaptation, Learning: The Science of Learning Models from Data*. Springer, 1996.

[88] F. Topsøe, Information theoretical optimization techniques. Kybernetika 15, 1979, 8-17.

[89] J. Uffink, Boltzmann’s Work in Statistical Physics, Stanford Encyclopedia Of Philosophy, 2004.

[90] A. Wakolbinger, Schrödinger Bridges from 1931 to 1991, in: E. Cabaña et al. (eds) , *Proc. of the 4th Latin American Congress in Probability and Mathematical Statistics*, Mexico City 1990, Contribuciones en probabilidad y estadistica matematica 3 (1992) , pp. 61-79.

[91] S. Yu and P.Mehta, The Kullback-Leibler rate pseudo-metric for comparing dynamical systems, *IEEE Trans. Automatic Control*, vol. 55, no. 7, pp. 15851598, 2010.