Peripheral Arterial Disease and Risk of Atrial Fibrillation and Stroke: The Multi-Ethnic Study of Atherosclerosis

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**Background**—Peripheral arterial disease (PAD) shares several risk factors with atrial fibrillation (AF), and persons with PAD have an increased risk of stroke. It is unclear if PAD is associated with an increased risk for AF and whether this potential association explains the increased risk of stroke observed in those with PAD.

**Methods and Results**—We examined the association between PAD, measured by ankle-brachial index (ABI), and incident AF and incident stroke, separately, in 6568 participants (mean age 62 ± 10 years, 53% women, 62% nonwhite) from the Multi-Ethnic Study of Atherosclerosis (MESA). ABI values <1.0 or >1.4 defined PAD. AF was ascertained through review of hospital discharge records and from Medicare claims data until December 31, 2010. An independent adjudication committee ascertained stroke events. Cox regression was used to estimate hazard ratios and 95% CIs for the association between PAD and AF and stroke. Over a median follow-up of 8.5 years, 301 (4.6%) participants developed AF and 140 (2.1%) developed stroke. In a model adjusted for sociodemographics, cardiovascular risk factors, and potential confounders, PAD was associated with an increased risk of AF (hazard ratio 1.5, 95% CI 1.1 to 2.0). In a similar model, PAD was associated with incident stroke (hazard ratio 1.7, 95% CI 1.1 to 2.5), and the magnitude of risk was not different after inclusion of AF as a time-dependent covariate (hazard ratio 1.7, 95% CI 1.1 to 2.5).

**Conclusions**—PAD is associated with an increased risk of AF and stroke in MESA. Potentially, the relationship between PAD and stroke is not mediated by AF. (J Am Heart Assoc. 2014;3:e001270 doi: 10.1161/JAHA.114.001270)

**Key Words:** ankle-brachial index • atrial fibrillation • peripheral arterial disease • stroke
Methods

Study Population

Details of MESA have been reported previously. Briefly, between July 2000 and September 2002, 6814 participants were recruited at 6 field centers (Baltimore, MD; Chicago, IL; Forsyth County, NC; Los Angeles, CA; New York, NY; and St. Paul, MN) in the United States. The institutional review board at each site approved the study and each participant gave informed consent at the time of enrollment. Requirement for study participation included age between 45 and 84 years and no history of clinical cardiovascular disease. For the purpose of this analysis, participants were excluded if they did not undergo baseline ABI measurement, baseline characteristics were not recorded, and/or follow-up data regarding AF or stroke were missing. Additionally, baseline cases of AF were excluded to examine incident AF.

Of the 6814 participants enrolled in MESA, we excluded 58 participants who had a diagnosis of AF before enrollment. The majority of these cases were detected by Medicare linkage, and only 1 case was present in the baseline electrocardiogram. Of those who remained, 6 participants with missing follow-up data, 77 participants with missing ABI data, and 105 participants with either missing baseline characteristics or missing medication data also were excluded. After all exclusions (n=246), 6568 participants (mean age 62±10 years, 53% women, 38% whites, 27% blacks, 22% Hispanics, and 12% Chinese Americans) remained and were included in the final analysis.

Baseline Characteristics

Participant characteristics recorded during the initial MESA visit were used in this analysis. Age, sex, race/ethnicity, income, and education were self-reported. Annual income was categorized into 3 levels (<$20 000, $20 000 to $49 999, and ≥$50 000). Similarly, education was categorized into “high school or less,” “some college,” and “college or more.” Smoking was defined as current or former smoker. Blood samples were obtained after a 12-hour fast. Serum measurements of total cholesterol, high-density lipoprotein (HDL) cholesterol, plasma glucose, and high-sensitivity C-reactive protein (hs-CRP) were used in this analysis. Diabetes was defined as a fasting glucose value ≥126 mg/dL or a history of diabetes medication use. After the participant rested for 5 minutes in a seated position, blood pressure was measured 3 separate times and the mean of the last 2 values was used. The use of aspirin, statins, and antihypertensive and lipid-lowering medications was self-reported. Body mass index was computed as the weight in kilograms divided by the square of the height in meters. Left ventricular hypertrophy was defined by the Cornell criteria (R-wave amplitude AV, plus S-wave amplitude V, ≥2800 mm for men and ≥2000 mm for women) using electrocardiographic data.

Peripheral Arterial Disease

Systolic blood pressures were measured in both the right and left brachial, posterior tibial, and dorsalis pedis arteries with Doppler instruments. The average of the measurements was used to calculate ABI for each side. Consistent with current guidelines, a value of <1.0 or >1.4 defined PAD. In a secondary analysis, we examined the association of low (<1.0) and high (>1.4) ABI values (referent group 1.0≤ABI≤1.4) separately with AF. Additionally, the association of ABI as a continuous variable with AF was examined and is presented as the risk of AF with decreasing ABI values by decrements of 0.1 after excluding participants with ABI values >1.4.

Atrial Fibrillation

Follow-up phone calls to study participants every 9 to 12 months were used to identify hospitalizations and medical records, including discharge diagnoses. Additionally, for participants 65 years or older enrolled in fee-for-service Medicare, Medicare claims data were used to identify inpatient AF cases. Incident AF was defined by International Classification of Disease, Ninth Revision codes 427.31 or 427.32.

Stroke

Trained personnel abstracted medical records to identify possible stroke events. Two physicians from the MESA Events Committee independently reviewed medical records for end point classification and assignment of event dates. If the reviewing physicians disagreed on the event classification, they adjudicated differences. If disagreements persisted, the full Events Committee made the final classification. Stroke was classified as a dichotomous variable (present or absent) and was defined by the rapid onset of documented focal neurologic deficits lasting ≥24 hours or until death. If symptoms were present for <24 hours, a clinically relevant lesion on brain imaging must have been present for stroke to be documented as present. Due to the small number of documented cases of stroke (n=140), we grouped hemorrhages (eg, subarachnoid, intraparenchymal) and infarcts into 1 category.

Statistical Analysis

Categorical variables were reported as frequency and percentage, while continuous variables were recorded as mean±standard deviation (SD). Statistical significance for categorical variables was tested using the χ² method and
the Wilcoxon rank-sum procedure for continuous variables. Kaplan–Meier estimates were used to compute cumulative incidence of AF by PAD, and the difference in estimates was compared by using the log-rank procedure. Follow-up time was defined as the time between the initial study visit until the diagnosis of AF, loss to follow-up, death, or end of follow-up, which was December 31, 2010. Cox proportional hazards regression was used to compute hazard ratios (HRs) and 95% confidence intervals (CI) for the association between PAD and AF. Multivariable models were constructed with incremental adjustments as follows: Model 1 adjusted for age, sex, race/ethnicity, income, and education; Model 2 adjusted for Model 1 covariates plus smoking status, systolic blood pressure, diabetes, body mass index, total cholesterol, HDL cholesterol, aspirin, antihypertensive and lipid-lowering medications, hs-CRP levels, and left ventricular hypertrophy. Using similar models, we also examined the association between PAD and stroke with and without AF to explore the attenuating effect of AF on stroke among participants with PAD by adjusting for AF as a time-dependent covariate. Although there were a small number of stroke events (n=140), ischemic strokes represented the majority (n=116, 83%) and a sensitivity analysis was performed examining the mediation of AF on the association of PAD with ischemic strokes. We tested for interactions between our main effect variable and age, sex, and race/ethnicity. We also constructed a restricted cubic spline model to examine the graphical relationship between ABI and the risk for AF and stroke, and incorporated knots at the 5th, 50th, and 95th percentiles. The proportional hazards assumption was not violated in our analysis. Statistical significance was defined as P<0.05. SAS Version 9.3 (SAS Institute Inc) was used for all analyses.

Results
A total of 774 (12%) study participants had baseline PAD. Baseline characteristics for study participants stratified by PAD status are shown in Table 1. Participants with PAD were more likely to be older and female, to have a high school diploma or less, and to have an annual income < $20,000 than participants without PAD. A higher percentage of black participants (17%) had PAD than other races (11% whites, 7.8% Hispanics, 8.1% Chinese Americans; P<0.0001). PAD participants also were more likely to smoke, to have diabetes, and to have higher values for systolic blood pressure, HDL cholesterol, and hs-CRP. Additionally, participants with PAD were more likely to report using statins and antihypertensive and lipid-lowering medications.

During a median follow-up of 8.5 years, 301 (4.6%) participants developed AF. AF incidence rates per 1000 person-years were 12 (95% CI 9.5 to 15.4) for participants with PAD compared with 5.3 (95% CI 4.6 to 6.0) for participants without PAD. Figure 1 shows the unadjusted cumulative incidence curves for AF by PAD (log-rank P<0.0001).

Participants with PAD had twice the risk of AF in an unadjusted Cox regression model (HR 2.3, 95% CI 1.8 to 3.0). The HR remained significant after adjustment for sociodemographics, cardiovascular risk factors, and potential confounders, and the results were consistent across subgroups stratified by age, sex, and race/ethnicity (Table 2). Expanding the demographic model to include field center and expanding the smoking variable to include pack-years did not change the result (HR 1.5, 95% CI 1.1 to 2.0).

The associations between AF and high (>1.4) and low (<1.0) ABI values were examined separately and were in the same direction as the main result for PAD (ABI <1.0, adjusted HR 1.5, 95% CI 1.1 to 2.0; ABI >1.4, adjusted HR 1.8, 95% CI 0.65 to 4.8). The result for ABI values >1.4 was not significant due to the small number of participants in this group (n=40) and consequently a small number of AF cases (n=4). The HR for ABI as a continuous variable, decreasing by 0.1, was 1.1 (95% CI 1.04 to 1.2) after excluding participants with ABI values >1.4.

Incident stroke occurred in 140 (2.1%) participants. Only 36 (26%) cases of stroke had PAD, and of these, only 1 (2.8%) had a diagnosis of AF before stroke. Table 3 shows the HR for stroke among participants with PAD compared with those without PAD. Additionally, HRs are shown with and without AF as a time-dependent covariate. In a sociodemographically adjusted model, PAD was associated with twice the risk of stroke compared with no PAD. The magnitude of the association did not substantively change with the addition of AF. A similar result was observed when the association between PAD and stroke was examined in a model further adjusted for cardiovascular risk factors and potential confounders (Table 3). A sensitivity analysis including only ischemic strokes showed a trend for the association between PAD and stroke (HR 1.5, P=0.071). The addition of AF to this model did not substantively alter our findings (HR 1.5, P=0.060).

Figure 2 shows the association between ABI values and incident AF and stroke using a restricted cubic spline model. As shown, there was an increased risk for AF and stroke with decreasing ABI values. A similar trend was observed for increasing ABI values.

Discussion
In this analysis from MESA, we examined the association between PAD and incident AF and stroke, separately. We also...
examined whether a potential association between PAD and AF explained the known association between PAD and stroke. Our study revealed 2 key findings. First, PAD is associated with incident AF regardless of age, sex, race/ethnicity, and cardiovascular risk factors. Second, our results suggest that the known association between PAD and stroke, which also was confirmed in our study, is not mediated by AF.

A recent examination of data from the Women’s Health Initiative Study has shown that PAD, which was ascertained by self-reported history, is independently associated with incident AF among postmenopausal women (HR 1.53, 95% CI 1.37 to 1.72). This result is similar to that observed in the current study. Additionally, the racially and ethnically diverse population of men and women in MESA increases the generalizability of our results compared with the Women’s Health Initiative Study. Also, we used ABI to define PAD instead of self-reported history, which was used in the Women’s Health Initiative Study. Considering that a small number of patients with abnormal ABI values have symptoms of PAD, our results provide a more accurate estimate of the risk for AF associated with PAD.

Several reports have described the association of PAD with stroke. Our results confirm these findings in MESA by using ABI measurements. However, to our knowledge, no studies have examined the effect of incident AF on the association between PAD and stroke. Due to many shared risk factors, we hypothesized that the risk of stroke among those with PAD was mediated by AF. Our results do not support this claim. This provides useful information to better understand the underlying mechanisms which are needed to appropriately

### Table 1. Baseline Characteristics of Study Participants Stratified by PAD (N=6568)

| Characteristic                           | PAD (n=774) | No PAD (n=5794) | P Value* |
|-----------------------------------------|-------------|-----------------|----------|
| Age, mean (SD), years                   | 67 (10)     | 61 (10)         | <0.0001  |
| Male sex, %                             | 282 (36)    | 2832 (49)       | <0.0001  |
| Race/ethnicity, %                       |             |                 |          |
| White                                   | 282 (36)    | 2237 (39)       | <0.0001  |
| Black                                   | 314 (41)    | 1487 (26)       |          |
| Chinese American                        | 64 (8.3)    | 725 (13)        |          |
| Hispanic                                | 114 (15)    | 1345 (23)       |          |
| Education, %                            |             |                 |          |
| High school or less                     | 337 (44)    | 2051 (35)       |          |
| Some college                            | 248 (32)    | 1623 (28)       | <0.0001  |
| College or more                         | 189 (24)    | 2120 (37)       |          |
| Annual income, %                        |             |                 |          |
| <$20 000                                 | 280 (36)    | 1471 (25)       |          |
| $20 000 to $49 999                      | 275 (36)    | 2028 (35)       | <0.0001  |
| ≥$50 000                                | 219 (28)    | 2295 (40)       |          |
| Body mass index, mean (SD) kg/m²        | 28 (5.9)    | 28 (5.4)        | 0.41     |
| Current or former smoker, %             | 445 (57)    | 2803 (48)       | <0.0001  |
| Diabetes, %                             | 157 (20)    | 756 (13)        | <0.0001  |
| Systolic blood pressure, mean (SD), mm Hg| 136 (25)    | 125 (21)        | <0.0001  |
| Total cholesterol, mean (SD), mg/dL      | 196 (38)    | 194 (35)        | 0.21     |
| HDL cholesterol, mean (SD), mg/dL        | 52 (16)     | 51 (15)         | 0.043    |
| Antihypertensive medications, %         | 386 (50)    | 2012 (35)       | <0.0001  |
| Statins, %                              | 166 (21)    | 793 (14)        | <0.0001  |
| Aspirin, %                              | 197 (25)    | 1349 (23)       | 0.18     |
| Lipid-lowering medications, %           | 181 (23)    | 865 (15)        | <0.0001  |
| hs-CRP, mean (SD), mg/L                 | 5.0 (7.3)   | 3.6 (5.7)       | <0.0001  |
| Left ventricular hypertrophy, %         | 31 (4.0)    | 216 (3.7)       | 0.70     |

HDL indicates high-density lipoprotein; hs-CRP, high-sensitivity C-reactive protein; PAD, peripheral arterial disease; SD, standard deviation.

*Statistical significance for continuous data was tested using Wilcoxon rank-sum procedure and for categorical data using the χ² test.
target preventative interventions to reduce the burden of stroke.

The association between PAD and AF is not completely understood. Both conditions share several risk factors and are associated with increased levels of inflammation and platelet-mediated thrombosis. It is plausible that the association of PAD with AF reflects an underlying state (eg, increased levels of inflammation, thrombosis, and coagulation) with a predisposition for both conditions to influence the development of the other. Further research is needed to elucidate these underlying mechanisms.

Patients with a high burden of medial arterial calcification have high ABI values, which represents a different pathological state compared with low ABI values. In our study, the association of both low and high ABI values with AF was in the same direction as the main model for PAD. This indicates that PAD, regardless of the pathophysiological basis or pattern, is associated with an increased risk for the development of AF. Among diabetic patients, arterial stenosis and incompressibility may have resulted in the misclassification of PAD. Diabetes is present in ≈41% of patients with PAD, and the risk for AF is increased by 40% in these patients. Diabetes also has been associated with a 10% higher risk of subclavian stenosis and a greater prevalence of subclinical episodic AF. While a differential bias for AF cannot be ruled out, the direction likely is toward the null. Furthermore, our multivariable results were not substantively affected when controlling for diabetes.

Several limitations of the current analysis should be addressed. Incident AF cases were ascertained from hospitalization discharge records and inpatient Medicare claims data using International Classification of Disease, Ninth Revision codes, which possibly resulted in misclassification. However, this method has been shown to have adequate positive predictive value for the identification of AF events. Additionally, paroxysmal cases of AF potentially were missed due to its time-dependent nature. This would only make our results conservative because these missed AF cases would be included in the non-AF group. Also, the actual onset of AF cannot be precisely known, and cases may have occurred before hospital admission in which the diagnosis was confirmed. Routine monitoring for AF was not performed in MESA and therefore asymptomatic cases possibly were missed. However, we do not know of any reason to suggest that the resulting bias, if any, would have been differential in

**Table 2. Risk of AF With PAD**

| Subgroup          | Events/No. at Risk | Model 1* HR (95% CI) | P Value | Model 2† HR (95% CI) | P Value | Interaction P Value‡ |
|-------------------|--------------------|----------------------|---------|----------------------|---------|----------------------|
| All               | 301/6568           | 1.7 (1.3 to 2.2)     | 0.0004  | 1.5 (1.1 to 2.0)     | 0.0041  | —                    |
| Age, years§       |                    |                      |         |                      |         |                      |
| <62               | 43/3218            | 2.2 (0.93 to 5.4)    | 0.071   | 1.8 (0.74 to 4.4)    | 0.20    | 0.79                 |
| ≥62               | 258/3350           | 1.9 (1.4 to 2.5)     | <0.0001 | 1.7 (1.3 to 2.3)     | 0.0003  |                      |
| Sex               |                    |                      |         |                      |         |                      |
| Female            | 115/3454           | 1.7 (1.1 to 2.6)     | 0.0098  | 1.6 (1.1 to 2.5)     | 0.028   | 0.64                 |
| Male              | 186/3114           | 1.6 (1.1 to 2.4)     | 0.012   | 1.5 (1.01 to 2.2)    | 0.043   |                      |
| Race/ethnicity    |                    |                      |         |                      |         |                      |
| White             | 166/2519           | 1.6 (1.1 to 2.3)     | 0.022   | 1.5 (0.98 to 2.2)    | 0.064   | 0.71                 |
| Nonwhite          | 135/4049           | 1.8 (1.2 to 2.7)     | 0.0038  | 1.6 (1.04 to 2.4)    | 0.034   |                      |

AF indicates atrial fibrillation; CI, confidence interval; HR, hazard ratio; PAD, peripheral arterial disease.

*Adjusted for Model 1 covariates plus smoking status, systolic blood pressure, diabetes, body mass index, total cholesterol, high-density lipoprotein cholesterol, aspirin, antihypertensive and lipid-lowering medications, high-sensitivity C-reactive protein levels, and left ventricular hypertrophy.

†Interactions tested using Model 2.

‡Dichotomized at the median age for study participants.
nature, rather than merely reducing effect estimates toward the null. PAD was defined by ABI values <1.0 and >1.4 according to current guidelines. The association between PAD and ABI may vary depending on the definition of abnormal ABI (eg, <0.9). Hence, comparing our results with other studies that use different definitions of abnormal ABI should be done with caution. ABI values used in this analysis were obtained during a single visit, and results may vary with interval measurements. Differences in the measurement of ABI may have posed another source of variability. However, the measurements of ABI in the United States and in MESA are standardized. Furthermore, the small number of AF cases that occurred before stroke may have limited the power to detect the mediating effect of AF on stroke. Similar to other observational studies, we also acknowledge that residual confounding remains a possibility.

In conclusion, we have shown that PAD, as defined by abnormal ABI values, is associated with an increased risk of AF in MESA. Additionally, our results suggest that the relationship between PAD and stroke is not mediated by AF.

**Table 3.** Association of PAD With Risk of Stroke With and Without Adjustment for AF

|                | Events/No. at Risk | Model 1* HR (95% CI) | P Value | Model 1† With AF HR (95% CI) | P Value | Model 2‡ HR (95% CI) | P Value | Model 2‡† With AF HR (95% CI) | P Value |
|----------------|--------------------|----------------------|---------|-----------------------------|---------|----------------------|---------|-----------------------------|---------|
| No PAD         | 104/5794           | 1.0                  | —       | 1.0                         | —       | 1.0                  | —       | 1.0                         | —       |
| PAD            | 36/774             | 2.0 (1.4 to 3.0)     | 0.0004  | 2.1 (1.4 to 3.0)            | 0.0003  | 1.7 (1.1 to 2.5)     | 0.015   | 1.7 (1.1 to 2.5)            | 0.011   |

AF indicates atrial fibrillation; CI, confidence interval; HR, hazard ratio; PAD, peripheral arterial disease.

*Adjusted for age, sex, race/ethnicity, income, and education.

†AF adjusted as a time-dependent covariate.

‡Adjusted for Model 1 covariates plus smoking status, systolic blood pressure, diabetes, body mass index, total cholesterol, high-density lipoprotein cholesterol, aspirin, antihypertensive and lipid-lowering medications, high-sensitivity C-reactive protein levels, and left ventricular hypertrophy.

**Figure 2.** Risk of AF (A) and stroke (B) with ABI. Each hazard ratio was computed with the median ABI value of 1.12 as the referent group and was adjusted for age, sex, and race/ethnicity. Dotted lines represent the 95% CI. The hazard ratio for AF is depicted in (A). The hazard ratio for stroke is depicted in (B). ABI indicates ankle-brachial index; AF, atrial fibrillation; CI, confidence interval.
Further research is needed to examine the usefulness of ABI to identify patients who have an increased risk for the development of AF.

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