Abstract—In this paper we propose a new parameter estimator that ensures global exponential convergence of linear regression models requiring only the necessary assumption of identifiability of the regression equation, which we show is equivalent to interval excitation of the regressor vector. Continuous and discrete-time versions of the estimators are given. An extension to—separable and monotonic—non-linear parameterizations is also given. The estimators are shown to be robust to additive measurement noise and—not necessarily slow—parameter variations. Moreover, a version of the continuous-time estimator that rejects sinusoidal disturbances with unknown internal model is given. The estimator is shown to be applicable to the classical model reference adaptive control problem relaxing the conspicuous assumption of known sign of the high-frequency gain. Simulation results that illustrate the performance of the estimator are given.

Index Terms—Parameter estimation; identifiability; adaptive control; robustness; interval excitation

I. INTRODUCTION

The tasks of control, identification or prediction of the dynamics of an unknown nonlinear system is usually accomplished assuming that there exists an approximation to the true dynamics with a fixed vector of parameters that globally fits the dynamics. A typical scenario is to assume the dynamics is described by an ordinary differential (or difference) equation with unknown parameters, which are then estimated designing an on-line parameter estimator. In its simplest formulation it is assumed that these parameters enter linearly in the dynamic model leading to a relationship of the form \( y = \phi^\top \theta \), with \( y \in \mathbb{R}, \phi \in \mathbb{R}^q \) measurable signals and \( \theta \in \mathbb{R}^p \) a constant vector of unknown parameters—that we call in the sequel linear regression equation (LRE) with \( \phi \) the regressor vector. LREs associated with many control problems, including system identification [30], adaptive control [19], [32], [52], filtering and prediction [16], reinforcement learning [28] and sparse regression analysis [10], have been reported in the literature.

Very often, gradient descent-based or least-squares parameter adaptation algorithms are used to generate on-line estimates of the unknown parameters. This leads to a linear time-varying (LTV) dynamical system that describes the behavior of the estimation errors, called parameter error equations (PEE), that have been extensively studied in the literature. A fundamental result is that a necessary and sufficient condition for the global exponential stability (GES) of the PEEs is that the regressor vector satisfies a persistency of excitation (PE) condition—which is a uniform complete observability property for the associated LTV system [19], [32], [52]. Here we underscore the qualifier “exponential” because it is widely accepted that without this property parameter convergence cannot be ensured and the robustness of the schemes is seriously damaged. Moreover, without PE the transient performance of the estimators is highly unpredictable and only a weak monotonicity property of the estimation errors norm can be guaranteed.

This fragile situation is particularly clear in model reference adaptive control (MRAC), for which it has vividly been shown in [18], [50] that several instability mechanisms are present. In spite of intensive research efforts [19], [32] the various fixes that have been introduced in the estimators—that include projections, deadzones and integrator leakages—have only partially alleviated this problem. Indeed, as shown in [23], it has only been possible to establish a “continuity” property with respect to unmodeled dynamics and the preservation of signal boundedness in the face of noise. More precisely, it has been proven that there exists a sufficiently small bound on the norm of the error dynamics such that (some kind of) stability is preserved—but, unfortunately, this bound is not quantifiable from the data of the problem. Regarding the presence of noise or parameter variations only signal boundedness, again with no uniform bound, is established. It should be furthermore added that all of this “robustified” schemes rely on the introduction of the dynamic normalization introduced in [14]—and its importance for robustness established in [35], [48]—that, as thoroughly discussed in [31], slows-down the adaptation bringing some additional robustness problems.

Unfortunately, the PE property—which imposes a "spanning
behavior” to the signals—is rarely satisfied in applications, where the task is often to drive the signals to some constant value. Although it has recently been shown that (non-uniform) global asymptotic stability can be ensured under weaker assumptions [4], [49], these conditions are still extremely stringent for applications. Hence the interest to propose new adaptation algorithms that ensure, via GES, parameter convergence without PE. This research line has been intensively pursued in the last few years and some recent adaptive schemes, where the PE assumption is obviated, have been reported in the literature—see [39] for a recent survey.

To the best of the authors’ knowledge the first estimators where parameter convergence is guaranteed under the extremely weak assumption of interval excitation (IE) [27]—called initial excitation in [46] and excitation over a finite interval in [53]—are the concurrent and the composite learning schemes reported in [11] and [45], respectively. These algorithms, which incorporate the monitoring of past data to build a stack of suitable regressor vectors, are closer in spirit to off-line estimators. See also [25], [34] for two early references where a similar idea is explored. As is well-known, the main drawback of off-line estimators is their inability to track parameter variations, which is very often the main objective in applications. This situation motivates the interest to develop bona-fide on-line estimators that preserve the scheme’s alertness.

Parameter convergence under the IE assumption was also recently established for the scheme reported in [15], that has the additional feature of ensuring convergence in finite-time—see also [42, Propositions 6 and 7] and [44]. A potential drawback of this algorithm is that it critically relies on the inclusion of a dynamic extension that mimicks the dynamics of the PEE, which may adversely affect the robustness of the estimator, [42, Remark 7] and [41]. A similar difficulty is present in the algorithm recently proposed in [56].

In this paper we are interested in the solution to the following key problems (KP).

**KP1** Design an on-line estimator that ensures GES of the PEE under the weakest assumption that the LRE is identifiable.

**KP2** Prove that, with a slight variation of the estimator that solves KP1, it is possible to prove the following important features:

- **Robustness**—in a clear, quantifiable sense—to external disturbances and (not necessarily slow) parameter variations.
- **Rejection** of sinusoidal disturbances with unknown internal model.
- **Applicability** to a well-defined class of nonlinearly parameterized regressor equations (NLPRE).

The first solution to KP1 was given in the recent paper [24] with the standing assumption that the regressor $\phi$ is IE, which we show in this paper is equivalent to identifiability of the LRE. Instrumental for the development of the adaptation algorithm proposed in [24] are the following steps.

**S1** The use of the dynamic regressor extension and mixing (DREM) parameter estimation procedure, which was first proposed in [2] for continuous-time (CT) and in [6] for discrete-time (DT) systems. The construction of DREM estimators proceeds in two steps, first, the inclusion of a free, linear operator that creates an extended matrix LRE similar to the ones designed in [26], [29]. Second, a nonlinear manipulation of the data that allows to generate, out of an $q$-dimensional LRE, $q$ scalar, and independent, LREs. DREM estimators have been successfully applied in a variety of identification and adaptive control problems, both, theoretical and practical ones, see [39], [42] for an account of some of these results.

**S2** The utilization of a procedure—proposed in [7]—to generate, from a scalar LRE, new scalar LREs where the new regressor satisfies some excitation conditions, even in the case when the original regressor is not exciting. To achieve this objective the authors borrow the key idea of the generalized parameter estimation based observer (GPEBO) [37], [43], to generate the new LRE that includes some free signals. Then, applying the energy pumping-and-damping injection principle of [58], these signals are selected to guarantee some excitation properties of the new regressor. Unfortunately, to prove in [7] that the aforementioned excitation properties guarantee GES it is necessary to assume some a priori non-verifiable conditions [7, Proposition 3]—in particular the absolute integrability of a signal and a non-standard requirement on the limiting behavior of some of the components of the trajectories of the estimator. Via the suitable selection of the aforementioned free signals in the new LRE, these two assumptions are relaxed in [24] providing a definite answer to KP1. Recalling the procedure followed in the construction of the estimator of [24], that is, first the application of DREM and then invoke GPEBO, we refer to it in the sequel as D+G. Interestingly, for the new estimator we also rely on the use of GPEBO and DREM, but under different circumstances and used in the opposite order, hence we refer to it in the sequel as G+D.

In this paper we provide an answer to the more challenging KP2, with our main contributions summarized as follows.

**C1** We prove, for the first time, that IE of the original LRE is equivalent to identifiability of the parameters. That is, to the existence of $q$ linearly independent regressor vectors for the reconstruction of an $q$-dimensional parameter vector.

**C2** The stability mechanisms and, consequently, the stability analysis of the G+D scheme is much more transparent than the ones of the D+G estimator. There are two consequences of this fact, on one hand, the procedure of tuning the estimator to achieve a satisfactory transient performance, which is difficult for the D+G scheme, is straightforward for the G+D one. On the other hand, by rendering the material accessible to a wider audience, the range of practical applicability of the new estimator is increased.
C3 The numerical complexity of the proposed estimator is considerably simpler than the D+G scheme. In particular, GPEBO is applied to the PEE of the classical gradient estimator avoiding the reference to the, rather obscure, concept of “virtual dynamics” used in the D+G estimator. Furthermore, the key mixing step of the DREM procedure reduces to a matrix multiplication, avoiding the need of generation of an extended LRE via the inclusion of additional LTV operators.

C4 The estimators are shown to be robust to additive measurement noise and—not necessarily slow—parameter variations. This feature is established showing that the estimator may be derived applying the DREM technique, which is the action of a linear operator on the original estimator. A variation of the CT estimator that rejects sinusoidal disturbances with unknown internal model is given. The qualifier “reject” in the present context means that it is possible to have a consistent estimate of the unknown parameters \( \theta \) in spite of the presence of the disturbances.

C5 The estimator is shown to be applicable to the classical MRAC problem, relaxing the conspicuous assumption of known sign of the high-frequency gain. As thoroughly discussed in [38, Subsection 1.2]—see also [5, Section 3]—this key assumption is hard to verify in practice, and the schemes that avoid it are, either technically unsound [19, Subsection 4.5.2] or only of theoretical interest, since their transient performance is intrinsically bad and practically inadmissible [33].

C6 Besides the case of LRE we consider (separable and monotonic) NLPRE, with the associated estimator preserving all the properties of the case of LRE.

C7 The behaviour of many physical systems is described via CT models. On the other hand, DT implementations of estimators are of significant practical relevance. Therefore, similarly to [24], [42], to comply with both scenarios we consider in the paper both kinds of LREs. Interestingly, in contrast to [24], the construction and analysis tools of both cases are essentially the same.

The remainder of the paper is organized as follows. In Section II we prove the equivalence between IE of the regressor and identifiability of the parameters of the LRE. Section III contains our main result for LRE. The proof that the proposed G+D estimator may be derived applying the DREM technique is given in Section IV. This important result is then used in Section V to carry-out the robustness analysis, including the proof of BIBO-stability and disturbance rejection. In section VI we apply the G+D estimator to relax the key assumption of known sign of the high-frequency gain in MRAC. In Section VII we extend the results for a class of NLPRE. Section VIII presents some simulation illustrating our main results. The paper is wrapped-up with concluding remarks and future research in Section IX. To simplify the reading, some preliminary lemmata are given in the Appendix and a list of acronyms is included at the end of the paper.

Notation. \( I_q \) is the \( n \times n \) identity matrix and \( 0_{n \times q} \) is an \( n \times q \) matrix of zeros. \( \mathbb{R}_{>0}, \mathbb{R}_{\geq 0}, \mathbb{Z}_{>0} \) and \( \mathbb{Z}_{\geq 0} \) denote the positive and non-negative real and integer numbers, respectively. For \( x \in \mathbb{R}^n \), we denote saure of the Euclidean norm as \( |x|^2 := x^\top x \). Given \( q \in \mathbb{Z}_{>0} \) we define the set \( \bar{q} := \{1, 2, \ldots, q\} \), CT signals \( s : \mathbb{R}_{\geq 0} \to \mathbb{R} \) are denoted \( s(t) \), while for DT sequences \( s : \mathbb{Z}_{>0} \to \mathbb{R} \) we use \( s(k) \). When a formula is applicable to CT signals and DT sequences the time argument is omitted. The symbol \( \| \cdot \|_\infty \) stands for the infinity norm of a signal or sequence. The action of an operator \( H \) on a CT signal \( s(t) \) is denoted as \( \mathcal{H}[s](t) \), and \( \mathcal{H}[s](k) \) for a sequence \( s(k) \). In particular, we define the derivative operator \( \mathcal{D}^n s(t) := \frac{d^n s(t)}{dt^n} \) and the delay operator \( q^n s(k) := s(k+n) \), where \( n \in \mathbb{Z}_{>0} \).

II. INTERVAL EXCITATION IS EQUIVALENT TO IDENTIFIABILITY

Throughout the paper we deal with LRE of the form

\[
y = \phi^\top \theta,
\]

where \( y \in \mathbb{R}, \phi \in \mathbb{R}^q \) are measurable signals and \( \theta \in \mathbb{R}^q \) is a constant vector of unknown parameters.\(^1\) The main objective of the paper is to provide a solution to KP2. To streamline the main result we need the following definition.

Definition 1: A bounded signal \( \phi \in \mathbb{R}^q \) is IE [27], [53] if

\[
\int_0^{t_c} \phi(s)\phi^\top(s)ds \geq C_c I_q,
\]

for some \( C_c \in \mathbb{R}_{>0} \) and \( t_c \in \mathbb{R}_{>0} \) in CT and

\[
\sum_{j=0}^{k_d} \phi(j)\phi^\top(j) \geq C_d I_q,
\]

for some \( C_d \in \mathbb{R}_{>0} \) and \( k_d \in \mathbb{Z}_{>0} \) in DT.

In this section we prove the fundamental result that IE of \( \phi \) is equivalent to identifiability of the LRE (1). We recall that identifiability, which is defined below, is a necessary and sufficient condition to reconstruct (even off-line) the unknown parameters.

Definition 2: The LRE (1) is said to be identifiable if and only if there exists a set of time instants—\( \{t_i\}_{i \in \bar{q}}, t_i \in \mathbb{R}_{>0} \) in CT and \( \{k_i\}_{i \in \bar{q}}, k_i \in \mathbb{Z}_{>0} \) in DT—such that

\[
\text{rank} \left\{ \left| \phi^\top(\tau_2) \right| \cdots \left| \phi^\top(\tau_q) \right| \right\} = q,
\]

where \( \tau_i = t_i \) in CT and \( \tau_i = k_i \) in DT.

Proposition 1: The LRE (1) is identifiable if and only if the regressor vector \( \phi \) is IE.

Proof: The proof of the DT version is obvious recalling that for any symmetric matrix \( A \in \mathbb{R}^{q \times q} \) we have the following equivalence

\[
A > 0 \iff z^\top Az > 0, \forall z \in \mathbb{R}^q \setminus \{0\}
\]

that, given the definition of IE, imposes the constraint \( k_d \geq q \).

\(^1\)To simplify the notation we consider the case of scalar \( y \), as will become clear below, the extension to the matrix case is straightforward.
The proof of the CT case proceeds as follows. The necessity is proved by contradiction. We suppose that there exists a positive integer \( q_0 < q \) such that
\[
\text{rank} \left[ \phi(t_1) \cdots \phi(t_q) \right] \leq q_0
\]
holds for all time sequence \( \{t_i\}_{i \in \bar{q}} \), with \( \{t_i\}_{i \in \bar{q}} \) being such that
\[
\text{rank} \left[ \phi(t_1) \cdots \phi(t_{q_0}) \right] = q_0.
\]
Let \( h \in \mathbb{R}^q \) be such that \( |h| = 1 \) and
\[
\phi^T(t_1)h = 0, \quad \forall i \in \bar{q}.
\]

Next we show that \( \phi^T(t)h = 0 \) for all \( t \in \mathbb{R}_{\geq 0} \) by contradiction. We suppose that there exists a \( t^* \in \mathbb{R}_{\geq 0} \) such that
\[
\phi^T(t^*)h \neq 0.
\]
This indicates
\[
q_0 = \text{rank} \left[ \phi(t_1) \cdots \phi(t_{q_0}) \right] < \text{rank} \left[ \phi(t_1) \cdots \phi(t_q) \right] \phi(t^*)
\]
which contradicts with the IE condition. Therefore, it can be seen that for any \( h \in \mathbb{R}^q \) satisfying \( |h| = 1 \), there always exists an \( \hat{i} \in \bar{q} \) such that
\[
|\phi(t_{\hat{i}})h| > 0.
\]

By continuity, it follows that for any \( h \in \mathbb{R}^q \) satisfying \( |h| = 1 \), there exists an \( \epsilon > 0 \) such that
\[
\sum_{i \in \bar{q}} |\phi(t_{\hat{i}})h| \geq |\phi(t_{\hat{i}})h| > 0, \quad \forall \hat{i} \in [t_1, t_1 + \epsilon]
\]
yielding
\[
\int_0^{t_1} |\phi^T(\tau)h|^2 d\tau \geq \int_0^{t_1 + \epsilon} |\phi^T(\tau)h|^2 d\tau > 0.
\]

Therefore, by recalling that such \( h \) is arbitrary, it can be concluded that \( \int_0^{t} \phi^T(\tau)h|^2 d\tau > 0 \) for \( t > t_1 \). The proof is thus completed.

Remark 1: For the sake of simplicity, we present \( y \) and \( \phi \) in (1) as functions of time, in the understanding that they may be functions of measurable signals evaluated at time \( t \) in CT or \( k \) in DT, for instance, the state, input and/or output of a dynamical system—see [39] and Sections VI and VIII for particular examples. Also, following standard practice in identification and adaptive control, in the sequel we disregard the presence of the exponentially decaying term stemming from the effect of the initial conditions of various filters used to generate the regression, see [1, Lemma 2] where the effect of this term in the DREM estimator is rigorously analyzed.

### III. MAIN RESULT FOR LINEAR REGRESSION EQUATIONS

In this section we present the G+D estimators that solve KP2 in CT and DT for the LRE (1).

**Proposition 2:** Consider the LRE (1). Define the \( G+D \) interlaced estimator
\[
\mathcal{S}_A[\hat{\theta}_g] = A\hat{\theta}_g + g\phi_y, \quad \hat{\theta}_g(0) = \theta_{g0} \in \mathbb{R}^q
\]
(3a)
\[
\mathcal{S}_b[\Phi] = \Phi, \quad \Phi(0) = I_q
\]
(3b)
\[
\mathcal{S}_a[\hat{\theta}] = \delta \Delta (Y - \Delta \hat{\theta}), \quad \hat{\theta}(0) = \theta_0 \in \mathbb{R}^q,
\]
(3c)
where the operators \( \mathcal{S}_a[\cdot] \) and \( \mathcal{S}_b[\cdot] \) are defined as
\[
\mathcal{S}_a[s] := \begin{cases} \mathcal{P}[s](t) & \text{in CT} \\ (q-1)[s](k) & \text{in DT} \end{cases},
\]
(4)
\[
\mathcal{S}_b[s] := \begin{cases} \mathcal{P}[s](t) & \text{in CT} \\ q[s](k) & \text{in DT} \end{cases},
\]
(5)
the functions \( g, \delta \) and \( \mathcal{A} \) are given by
\[
g := \begin{cases} \gamma_{g}(t) & \text{in CT} \\ \frac{1}{\gamma_{g}(k)+|\phi(k)|} & \text{in DT} \end{cases}, \quad \delta := \begin{cases} \gamma & \text{in CT} \\ \frac{1}{\gamma+\Delta_{g}(k)} & \text{in DT} \end{cases}
\]
\[
\mathcal{A} := \begin{cases} -g(t)\phi(t)\phi^T(t) & \text{in CT} \\ I_q - g(k)\phi(k)\phi^T(k) & \text{in DT} \end{cases}
\]
(6)
with \( \gamma_{g}(\cdot) \in \mathbb{R}_{>0}, \gamma \in \mathbb{R}_{>0}, \) and the functions
\[
\mathcal{D} := I_q - \Phi
\]
(7a)
\[
\Delta := \text{det}(\mathcal{D})
\]
(7b)
\[
Y := \text{adj}(\mathcal{D})(\hat{\theta}_g - \Phi\theta_{g0}),
\]
(7c)
where \( \text{adj}(\cdot) \) stands for the adjugate matrix. If the LRE (1) is identifiable then
\[
\lim_{\tau \to \infty} \hat{\theta}(\tau) = 0, \quad (exp)
\]

\text{Proof:} Replacing (1) in (3a) yields the PEE for the gradient estimator
\[
\mathcal{S}_a[\hat{\theta}_g] = A\hat{\theta}_g,
\]
where \( \hat{\theta}_g := \hat{\theta} - \theta \) and we used the definition (6). Consequently, from the properties of the fundamental matrix \( \Phi \) [51] defined in (3b), we get
\[
\hat{\theta}_g = \Phi\hat{\theta}_g(0),
\]
(8)
which may be rewritten as the extended LRE
\[
\mathcal{D}\hat{\theta} = \hat{\theta}_g - \Phi\theta_{g0},
\]
(9)
where we used (7a). Following the DREM procedure we multiply (9) by \( \text{adj}(D) \) to get the following scalar LRE

\[
Y_t = \Delta \theta_t, \quad i \in \bar{q},
\]

where we used (7b) and (7c). We underscore the fact that the regressor \( \Delta \) is a scalar.

Replacing (10) in (3c) yields the PEE for each of the elements \( \hat{\theta}_i \), \( i \in \bar{q} \), of the vector \( \hat{\theta} \) of the least mean squares estimator (3c)

\[
\delta_{\theta_i}[\hat{\theta}_i] = -\phi \Delta^2 \hat{\theta}_i.
\]

Now, in Proposition 1 it is shown that identifiability is equivalent to \( \phi \) in IE. On the other hand, in Lemmas 3 and 5, given in Appendix A, we prove that the IE assumption implies that \( \Delta \) is PE in the CT and DT case, respectively. The proof of exponential convergence in CT follows from the well-known result [52, Theorem 2.5.1].

For the DT case we have the following argument. The PEEs for the normalized least mean squares estimator (3c) are given by

\[
\hat{\theta}_i(k + 1) = \hat{\theta}_i(k) - \frac{\Delta^2(k)}{\gamma + \Delta^2(k)} \hat{\theta}_i(k)
\]

\[
= \frac{\gamma}{\gamma + \Delta^2(k)} \hat{\theta}_i(k)
\]

\[
= \prod_{j=0}^{\infty} \frac{\gamma}{\gamma + \Delta^2(j)} \hat{\theta}_i(0).
\]

From the fact that \( \Delta(k) \in \mathbb{R}_{\geq 0} \) for all \( k \geq k_d \) we conclude that \( \Delta(k) \notin \ell_2 \). This, together with the fact that

\[
\Delta(k) \notin \ell_2 \Leftrightarrow \prod_{j=0}^{\infty} \frac{\gamma}{\gamma + \Delta^2(j)} = 0,
\]

proves global convergence. The proof that the convergence is exponential follows from the inequality

\[
\prod_{j=0}^{k} \frac{\gamma}{\gamma + \Delta^2(j)} \leq \exp \left( -\sum_{j=0}^{k} \frac{\Delta^2(j)}{\gamma + \Delta^2(j)} \right).
\]

Remark 2: It is important to underscore that the estimator of Proposition 1 consists of the interlacing of two classical gradient-based parameter search algorithms and contains only two tuning gains \( \gamma \) and \( \gamma_p \). The effect of both gains on the transient performance of the estimator is very clear and has been extensively studied in the literature—see [13] for a recent survey of the main results on this topic. The importance of this fact can hardly be underestimated because, as is well-known, the stage of commissioning the estimators, which is usually done with a trial-and-error approach, is very painful and a bad tuning has a serious deleterious effect on the overall performance of the scheme.

Remark 3: Notice that in the interval \( t \in [0, t_v) \) in CT or \( k \in [0, k_v) \) in DT the IE condition is not yet satisfied, which implies that \( \Delta = 0 \) in this interval. Consequently, the second estimator remains frozen in this interval, that is \( \hat{\theta} = \hat{\theta}_0 \). We can, therefore, interpret the role of both estimators as follows: the one of \( \hat{\theta}_g \) “gathers” the required excitation while the one of \( \hat{\theta} \) starts “operating” only after we have a rich regressor.

IV. A DREM PERSPECTIVE OF THE PROPOSED ESTIMATORS

In this section we show that the reparameterization (9) used in the estimator of Proposition 2 can be generated applying the well-known DREM procedure [2] to the LRE (1) in both, the CT and the DT cases. Towards this end, we recall that the first step in DREM is to generate an extended regressor applying a linear, single-input \( q \)-output operator \( H \) to the LRE (1). Because of linearity, this yields the new (extended) LRE

\[
y = \phi \theta
\]

where we defined the vector \( y \in \mathbb{R}^q \) and the matrix \( \phi \in \mathbb{R}^{q \times q} \) as

\[
y := H[y]
\]

\[
\phi := [H[\phi_1] | H[\phi_2] | \cdots | H[\phi_q]].
\]

In the proposition below we identify an LTV operator \( H \) such that

\[
y = \hat{\theta}_g - \Phi \theta y_0
\]

\[
\phi = D,
\]

yielding the extended LRE (9), hence proving the claim above.

Proposition 3: Define the single-input \( q \)-output operator \( H : u \mapsto u \), with \( u \in \mathbb{R} \), \( u \in \mathbb{R}^q \), and state-space representation

\[
\delta_{\theta_u}[x_u] = Ax_u + g \phi u
\]

\[
u = x_u,
\]

where the operator \( \delta_{\theta_u} \) and the functions \( A \) and \( g \) are defined in Proposition 2 and the initial condition of the state \( x_u \) is zero. Applying this operator to the signal of the LRE (1) we obtain the LTV systems

\[
\delta_{\theta_u}[x_q] = Ax_q + g \phi y
\]

\[
\delta_{\theta_u}[x_{\phi,i}] = Ax_{\phi,i} + g \phi \phi_i,
\]

with initial conditions \( x_q(0) = 0 \times q \times 1 \) and \( x_{\phi,i}(0) = 0 \times q \times 1 \), \( i \in \bar{q} \). Then, for all \( t \in \mathbb{R}_{\geq 0} \) in CT and all \( k \in \mathbb{Z}_{\geq 0} \) in DT, the following identity holds

\[
\hat{\theta}_g - \Phi \theta y_0 = x_y
\]

\[
D = [x_{\phi,1} | x_{\phi,2} | \cdots | x_{\phi,q}].
\]

Proof: First, notice that we have

\[
\delta_{\theta_u}[D] = \delta_{\theta_u}[I_q] - \delta_{\theta_u}[\Phi]
\]

\[
= \delta_{\theta_u}[I_q] - A \Phi
\]

\[
= \delta_{\theta_u}[I_q] - A[I_q - D]
\]

\[
= AD - A + \delta_{\theta_u}[I_q]
\]

\[
= AD + g \phi \phi^T
\]

where we have used (6) and the fact that \( \delta_{\theta_u}[I_q] = 0 \) in CT and \( \delta_{\theta_u}[I_q] = I_q \) in DT to derive the last equality.

Now, define the state vector errors

\[
\hat{x}_y := x_y - (\hat{\theta}_g - \Phi \theta y_0)
\]

\[
\hat{x}_{\phi,i} := x_{\phi,i} - D_i,
\]

\( i \in \bar{q} \).
with \( D \), the \( i \)-th column of \( D \) and notice that
\[
\begin{align*}
\hat{S}_n[x_y] &= \bar{A}x_y \\
\hat{S}_n[x_{\phi,i}] &= \bar{A}x_{\phi,i}, \quad i \in \bar{q}.
\end{align*}
\]

The proof is completed noting that \( \hat{x}_g(0) = 0_{q \times 1} \) and \( \hat{x}_{\phi,i}(0) = 0_{q \times 1} \), hence (15) holds true for all \( t \in \mathbb{R}_{\geq 0} \) in CT and all \( k \in \mathbb{Z}_{\geq 0} \) in DT.

**Remark 4:** In DREM it is usually assumed that the operator \( \mathcal{H} \) is bounded-input bounded-output (BIBO)-stable. This condition is imposed to preserve boundedness of the extended LRE (12), which however is not necessary for the overall stability analysis of the DREM estimator. It is possible to show that—without further assumptions on \( \gamma_g \), besides positivity—the operator \( \mathcal{H} \), defined via (13), is not BIBO-stable.

**V. Robustifying the Proposed Estimators**

In this section we analyze the robustness vis-à-vis additive perturbation of (a slight variation of) the estimator of Proposition 2. That is, we consider the perturbed LRE
\[
y = \phi^T \theta + d,
\]
where \( d \) represents an additive perturbation signal. This signal may come from additive noise in the measurements of \( y \) and \( \phi \) or time variations of the parameters, that is, \( d \) may be decomposed as
\[
d = d_y + d_{\theta}^T \theta + d_{\phi} \phi,
\]
where \( d_y \in \mathbb{R} \) and \( d_{\theta} \in \mathbb{R}^q \) represent the measurement noise added to \( y \) and \( \phi \), respectively, and \( d_{\phi} \in \mathbb{R}^q \) captures time variations in the parameters. We make the reasonable assumption that these signals are all bounded.

**A. BIBO stability**

In this subsection we prove that, imposing an additional condition on the adaptation gain \( \gamma_g \), it is possible to robustify the proposed estimators with respect to the additive disturbance \( d \). More precisely, we will prove the parameter estimation error \( \hat{\theta} \) remains bounded. Instrumental to establish these results are, on one hand the claim of GES of the unperturbed estimator of Proposition 2 and, on the other hand, the proof in Proposition 3 that the estimators can be derived from the DREM procedure. As indicated in Remark 3 the operator \( \mathcal{H} \) is defined with arbitrary, positive \( \gamma_g \) in (13)—is not BIBO-stable, hence in this subsection an additional constraint on \( \gamma_g \) is imposed to ensure that the operator \( \mathcal{H} \) is BIBO-stable. Consequently, the presence of the disturbance \( d \) will induce an additive bounded disturbance on the extended LRE (9), yielding a GES system with a bounded perturbation.

The main result is summarized in the proposition below.

**Proposition 4:** Consider the perturbed LRE (16) with bounded \( d \). Assume the unperturbed LRE (1) is identifiable. If the adaptation gain \( \gamma_g \in \mathbb{R}_{>0} \) is selected such that \( \int_0^\infty \gamma_g(t)dt \) is bounded in CT or \( \sum_{k=0}^{\infty} \frac{1}{\gamma_g(k)} \) is bounded in DT, the G+D estimator of Proposition 2 is robust in the sense that the parameter estimation error \( \hat{\theta} \) remains bounded.

\[
\begin{align*}
\text{Proof}: & \quad \text{Applying the operator } \mathcal{H} \text{ of Propositions 3 to the perturbed LRE yields the perturbed version of the extended LRE (9) as} \\
& \quad \hat{\theta}_g - \Phi \theta_0 = D \theta + \mathcal{H}[d],
\end{align*}
\]
where we exploited the property of linearity of \( \mathcal{H} \). Next we proceed to show that, under the conditions on \( \gamma_g \) imposed in the proposition, the operator \( \mathcal{H} \) is BIBO-stable. This is done by proving that, for all bounded \( d \), the signal \( \mathcal{H}[d] \) is also bounded.

For the CT case, the signal \( \mathcal{H}[d](t) \) is generated via the CT LTV system
\[
\dot{x}_d(t) = -\gamma_g(t)\phi(t)\phi^T(t)x_d(t) + \gamma_g(t)\phi(t)d(t)
\]
\[\mathcal{H}[d](t) = x_d(t).
\]
Defining \( V(x_d) := |x_d|^2 \), we have
\[
\dot{V} \leq -2\gamma_g(t)|\phi(t)x_d(t)|^2 + 2\gamma_g(t)|\phi^T(t)x_d(t)||d(t)|
\]
\[\leq -\gamma_g(t)|\phi^T(t)x_d(t)|^2 + \gamma_g(t)|d(t)|^2,
\]
which yields
\[
V(t) - V(0) \leq \int_0^t \gamma_g(\tau)|d(t)|^2d\tau
\]
\[\leq \|d(t)\|_2^2 \int_0^t \gamma_g(\tau)d\tau.
\]
This further implies that \( x_d(t) \), hence \( \mathcal{H}[d](t) \), is bounded as \( d(t) \) and \( \int_0^\infty \gamma_g(\tau)d\tau \) are bounded.

For the DT case, the signal \( \mathcal{H}[d](k) \) is generated via the DT LTV system
\[
x_{d}(k+1) = \left(I_q - \frac{\phi(k)\phi^T(k)}{\gamma_g(k) + |\phi(k)|^2}\right)x_d(k) + \frac{\phi(k)d(k)}{\gamma_g(k) + |\phi(k)|^2}.
\]
Similarly to the CT case, for the function \( V(x_d) \) we obtain
\[
V(k+1) - V(k) \leq 4 \frac{|d(k)|^2}{\gamma_g(k)}
\]
which yields
\[
V(k+1) - V(0) \leq \sum_{j=0}^{k} 4 \frac{|d(j)|^2}{\gamma_g(j)}
\]
\[\leq \|d(k)\|_2^2 \sum_{j=0}^{k} \frac{4}{\gamma_g(j)}.
\]
As \( d(k) \) and \( \sum_{k=0}^{\infty} \frac{1}{\gamma_g(k)} \) are bounded this implies that \( x_d(k) \) is bounded.

From the analysis above, we conclude that the operator \( \mathcal{H} \) is BIBO-stable and \( \mathcal{H}[d] \) is bounded. Then, multiplying (17) by \( \text{adj}(D) \) we get the following perturbed LRE
\[
Y = \Delta \theta + \xi,
\]
where we defined the signal
\[
\xi := \text{adj}(D) \mathcal{H}[d].
\]
We notice that, due to Proposition 3 and the BIBO-stability of $\mathcal{H}$, this signal is bounded. Replacing (19) in the estimator (3c) yields the scalar, perturbed PEEs

$$\hat{\theta}_i(t) = -\gamma \Delta^2(t) \hat{\theta}_i(t) + \gamma \Delta(t) \xi_i(t), \ i \in q,$$

in CT and

$$\hat{\theta}_i(k + 1) = \frac{\gamma}{\gamma + \Delta^2(k)} \hat{\theta}_i(k) + \frac{\gamma \Delta^2(k)}{\gamma + \Delta^2(k)} \xi_i(k), \ i \in q$$

in DT. Notice that in both cases $\Delta$ is PE by the Lemmata 3 and 5 in the Appendix and thus we are dealing with GES scalar systems with bounded additive perturbations. The proof of the claims follows then invoking standard arguments of GES systems with bounded additive perturbations [20], [21].

**Remark 5:** As indicated above the additional assumption on $\gamma_g$ is introduced to robustify the estimator of $\hat{\theta}_g$ in (3a) in the sense that $\hat{\theta}_g$ is bounded if the perturbation $d$ is bounded. We remark that this objective may be achieved with other robustified estimators methods, such as the dead-zone [36] and projection [52], [53].

**Remark 6:** It is important to remark that the boundedness constraint imposed on the adaptation gain $\gamma_g$ essentially imposes that the estimator of $\hat{\theta}_g$ loses its alertness properties. However, since we are dealing with an interlaced estimator that incorporates a second DREM-based stage, the overall scheme does not necessarily loses the alertness. In this respect, it would be more interesting to use, as suggested in Remark 5 above, other robustification methods that do not suffer from this drawback.

### B. Rejection of sinusoidal disturbances with unknown internal model for CT LRE

In this subsection we consider the CT perturbed LRE (18) under the assumption that the disturbances $\xi_i(t)$, $i \in q$, are sinusoidal signals of the form

$$\xi_i(t) = a_i \sin(\omega_i t + \psi_i), \ i \in q,$$

with unknown amplitudes $a_i \in \mathbb{R}_{>0}$, frequencies $\omega_i \in \mathbb{R}_{>0}$ and phase shifts $\psi_i \in \mathbb{R}$.

Recalling that $\Delta(t)$ is a scalar signal, (18) consists of a set of scalar perturbed LREs of the form

$$Y_i(t) = \Delta(t) \theta_i + \xi_i(t). \quad (20)$$

The main result is the proof that, with a suitable dynamic extension and a second application of the GPEBO approach to each scalar entry, it is possible to derive an unperturbed LRE for $\theta_i$ to which we can apply the standard gradient or DREM estimators to reconstruct $\theta_i$.

The first step in the design is to, fix a constant $\lambda \in \mathbb{R}_{>0}$, and apply the filter

$$F(\mathcal{P}) := \frac{\lambda^2}{(\mathcal{P} + \lambda)^2},$$

to each element of (18), yielding the scalar, perturbed LRE

$$Y_F(t) = \Delta_F(t) \theta + \xi_F(t), \quad (21)$$

where we defined the filtered signals $\gamma_F := F(\mathcal{P})(\cdot)$. Notice that since

$$\xi(t) = a_i \sin(\omega t + \psi),$$

we have that $\xi_F(t) = -\omega^2 \xi_F(t)$, from which we obtain the parameterization

$$\xi_F(t) = -\frac{1}{\omega^2} \xi_F(t),$$

$$= \theta \varphi(t) F(t) - \frac{1}{\omega^2} \bar{Y}_F(t),$$

$$=: \varphi(t) \mu,$$ \quad (22)

where we used (21) to get the second identity, the unknown parameter vector $\mu$ is given as

$$\mu := \text{col}\left(\frac{1}{\omega^2}, \frac{1}{\omega^2}\right),$$ \quad (23)

and we defined the measurable regressor

$$\varphi(t) := (\Delta_F(t), \bar{Y}_F(t)).$$ \quad (24)

Obviously, since $\xi_F$ is unmeasurable, we cannot use (22) for the estimation of $\mu$. It is, at this point, where we use again GPEBO to overcome this difficulty—as shown in the proposition below.

**Proposition 5:** Consider the scalar, perturbed LRE (21). Define the dynamic extension

$$\dot{z}(t) = -z(t) - Y_F(t), \ z(0) = 0 \quad (25a)$$

$$\dot{r}(t) = A \xi(t) r(t) + b \xi(t), \ r(0) = \text{col}(0, 0) \quad (25b)$$

$$\dot{\Omega}(t) = A \xi(t) \Omega(t) - e_2 \varphi^T(t), \ \Omega(0) = \Theta_{2x2} \quad (25c)$$

$$\Phi \xi(t) = A \xi(t) \Phi \xi(t), \ \Phi \xi(0) = I_2, \quad (25d)$$

where $\varphi(t)$ is given in (24), and we defined the matrices

$$A \xi(t) := \begin{bmatrix} 0 & \Delta_F(t) \\ -\Delta_F(t) & -1 \end{bmatrix},$$ \quad (26a)

$$b \xi(t) := [-\Delta_F(t) z(t)]^T, \quad (26b)$$

and $e_2 := \text{col}(0, 1)$. The following LRE holds

$$z(t) - r_2(t) = [\Phi \xi(t)]_{2,1}(t) \ \Omega_{2,1}(t) \ \Omega_{2,2}(t) \begin{bmatrix} \theta \\ \mu_1 \\ \mu_2 \end{bmatrix},$$ \quad (27)

where the unknown vector $\mu$ is defined in (23).

**Proof:** Similarly to [7] notice that, since $\theta$ is constant, we can write

$$\dot{\theta}(t) = \Delta_F(t) z(t) - \Delta_F(t) \xi(t),$$

while the $z$-dynamics (25a) may be written as

$$\dot{z}(t) = -z(t) - \Delta_F(t) \theta - \xi_F(t),$$

where we used (21). Combining these two equations define the “virtual” dynamical system

$$\dot{\chi}(t) = A \xi(t) \chi(t) - \begin{bmatrix} \Delta_F(t) z(t) \\ \xi_F(t) \end{bmatrix}$$

$$= A \xi(t) \chi(t) + b \xi(t) - e_2 \varphi^T(t) \mu$$
where the “state” is \( \chi(t) := \text{col}(\theta(t), z(t)) \) and we used (26a) in the first identity and (22), (24) and (26b) to get the second equation.

Define the signal
\[
e_\xi(t) := r(t) - \chi(t) + \Omega(t)\mu,
\]
which satisfies \( \dot{e}_\xi(t) = A_\xi(t)e_\xi(t) \). Consequently, in view of the definition of \( \Phi_\xi(t) \), we have that
\[
e_\xi(t) = \Phi_\xi(t)e_\xi(0) = \Phi_\xi(t) \begin{bmatrix} -\theta \\ 0 \end{bmatrix},
\]
where we took into account the choice of initial conditions in (25). From the equation above we have that \( e_\xi(t) = -(\Phi_\xi)_{2,1}(t)\theta \). Now, taking the second element from the definition of \( e_\xi(t) \) we get
\[
z(t) = r_2(t) - e_\xi(t) + \Omega_{2,1}(t)\mu_1 + \Omega_{2,2}(t)\mu_2 = r_2(t) + (\Phi_\xi)_{2,1}(t)\theta + \Omega_{2,1}(t)\mu_1 + \Omega_{2,2}(t)\mu_2 .
\]
This completes the proof.

Remark 7: It is important to note that the standing assumption in this subsection is that it is the LRE (18), which is the LRE generated by the G+D procedure, that is perturbed by a sinusoidal disturbance. It is not clear which kind of disturbances \( d(t) \) to the original LRE (16) will give rise to a sinusoidal signal \( \xi(t) \) since the relation between these two signal involves complicated operations, namely (19).

Remark 8: Replacing (23) in (27) we have that
\[
\text{col}(\theta, \mu_1, \mu_2) = \text{col} \left( \theta, \frac{\theta}{\omega^2}, \frac{1}{\omega^2} \right),
\]

hence there are only two unknown parameters—\( \theta \) and \( \omega \)—and the LRE is overparameterized. In Section VII we give a procedure to identify the parameters of a class of NLPRE with our G+D method without overparametrization, which turns out to be applicable to this example.

Remark 9: It is clear that it is possible to define an alternative form for the \( z \)-dynamics (25a), for instance, adding constants \( a \in \mathbb{R}_{>0} \) and \( b \in \mathbb{R} \) as
\[
\dot{z}(t) = -az(t) - bY_P(t),
\]
and consequently redefine the matrices (26). This modification would add some additional tuning gains to the algorithm without modifying the final result.

VI. APPLICATION TO CT MODEL REFERENCE ADAPTIVE CONTROL

In this section we show that, using the G+D estimator of Proposition 2 in the classical problem of MRAC of scalar CT LTI systems it is possible to remove the standard assumption of knowledge of the high frequency gain [19, 32, 52, 53]. More precisely, we consider a CT plant
\[
D(\mathcal{P})y_p(t) = k_pN(\mathcal{P})u_p(t),
\]
with \( D(\mathcal{P}) \) and \( N(\mathcal{P}) \) monic and coprime with unknown coefficients and \( k_p \in \mathbb{R} \) is the unknown high frequency gain.

We make the following standard assumptions regarding the plant.

A.1 \( N(\mathcal{P}) \) is a Hurwitz polynomial.

A.2 The plan order \( n_p \) and relative degree \( n_p - m_p \geq 1 \) are known.

The MRAC objective is to asymptotically drive to zero the tracking error
\[
e_T(t) = y_p(t) - y_m(t)
\]
where
\[
y_m(t) = \frac{k_m}{D_m(\mathcal{P})}r(t)
\]
with \( D_m(\mathcal{P}) = \sum_{i=0}^{n_p-m_p} a_i^m \mathcal{P}^i \) a designer-chosen monic, Hurwitz polynomial, \( k_m \in \mathbb{R} \) and \( r(t) \) is a bounded reference signal.

Instrumental for the proposed MRAC is the lemma below, which includes the classical direct control model reference parameterization and the input error parameterization given in [52, Subsection 3.3.1].

Lemma 1: Consider the plant (28) and the tracking error (29). There exists a vector \( \theta \in \mathbb{R}^{2n_p} \) such that
\[
e_T(t) = \frac{k_p}{D_m(\mathcal{P})}[u_p(t) - \theta^T \phi_{\text{IE}}(t)],
\]
where the vector \( \phi_{\text{IE}}(t) \in \mathbb{R}^{2n_p} \) is given by
\[
\phi_{\text{IE}}(t) = \frac{1}{\lambda(\mathcal{P})} \text{col} \left( u_p(t), \dot{u}_p(t), \ldots, u_p^{(n_p-2)}(t), y_p(t), \ldots, y_p^{(n_p-2)}(t), \lambda(\mathcal{P})y_p(t), \lambda(\mathcal{P})r(t) \right)
\]
with \( \lambda(\mathcal{P}) = \sum_{i=0}^{n_p-m_p} \lambda_i^2 \mathcal{P}^i \) a designer-chosen monic, Hurwitz polynomial. Moreover the vector \( \theta \) satisfies the following input-error LRE
\[
u_{\text{IE}}(t) = \theta^T \phi_{\text{IE}}(t)
\]
where
\[
\phi_{\text{IE}}(t) = \begin{bmatrix} \phi_N(t) \\ \frac{1}{\lambda_m} y_p(t) \end{bmatrix}
\]
\[
\phi_N(t) = \frac{1}{D_m(\mathcal{P})\lambda(\mathcal{P})} \text{col} \left( u_p(t), \dot{u}_p(t), \ldots, u_p^{(n_p-2)}(t), \ldots, y_p(t), \dot{y}_p(t), \ldots, y_p^{(n_p-2)}(t), \lambda(\mathcal{P})y_p(t) \right)
\]
\[
u_{\text{IE}}(t) := \frac{1}{D_m(\mathcal{P})} u_p(t).
\]

Motivated by (30), MRAC designs are completed proposing a controller of the form
\[
u_p(t) = \tilde{\theta}^T(t) \phi_{\text{IE}}(t),
\]
where \( \tilde{\theta}(t) \in \mathbb{R}^{2n_p} \) are the estimates of the parameters \( \theta \), which are generated via a parameter adaptation algorithm. As an immediate corollary of Proposition 2 we have that
\footnote{Notice that \( \phi_N(t) \) consists of the first 2\( n_p \) elements of \( \phi_{\text{IE}}(t) \) passed through the filter \( \frac{1}{\lambda_m^2} \).}
using the LRE (32) to generate these estimates with the G+D algorithm solves the MRAC problem requiring only the classical Assumptions A.1 and A.2 with no additional assumption on \(k_p\)—instead, we require identifiability of the LRE (32), which is a necessary assumption for reconstruction of \(\theta\).

For the sake of completeness we summarize this result in the following.

**Proposition 6:** Consider the plant (28) satisfying Assumptions A.1 and A.2 and LRE (32) with the G+D interlaced estimator

\[
\hat{\theta}_g(t) = -\gamma_g(t)\phi(t)[\phi_t^\top(t)\hat{\theta}_g(t) - u(t)],
\]

\[
\hat{\Phi}(t) = \gamma_g(t)\phi(t)\phi_t^\top(t)\Phi(t), \quad \Phi(0) = I_{2np},
\]

\[
\hat{\theta}(t) = \gamma\Delta(t)[Y(t) - \Delta(t)\hat{\theta}(t)], \quad \hat{\theta}(0) = \theta_0 \in \mathbb{R}^{2np},
\]

where \(\hat{\theta}_g(0) = \theta_{g0} \in \mathbb{R}^{2np}\) and \(\gamma_g(t) \in \mathbb{R}_{>0}\) and we defined the functions

\[
D(t) := I_{2np} - \Phi(t)
\]

\[
\Delta(t) := \text{det}(D(t))
\]

\[
Y(t) := \text{adj}(D(t))\{\hat{\theta}_g(t) - \Phi(t)\theta_{g0}\}.
\]

If the LRE (32) is identifiable then

\[
\lim_{t \to \infty} \hat{\theta}(t) = 0, \quad (exp).
\]

Consequently, applying the control (33) to the plant (28) the tracking error (30) verifies

\[
\lim_{t \to \infty} e_T(t) = 0, \quad (exp).
\]

**Remark 10:** In [52] an estimator that uses the input-error parameterization (32) and ensures global tracking is proposed. Unfortunately, this parameter algorithm includes a projection technique that requires, besides the knowledge of sign(\(k_p\)), and upper bound on |\(k_p\)|, which is essential for the proof. Moreover, it has recently been shown in [5] that, in the absence of the projection, input-error MRAC suffers from an instability mechanism that may give rise to unbounded trajectories—even in the simplest case of first-order plants with \(r(t) = 0\).

**Remark 11:** As shown in [5] the instability mechanism of input-error MRAC is related with the lack of excitation in the regressor. This difficulty is avoided with the G+D estimator that ensures the regressor—in this case \(\Delta(t)\)—is PE.

**Remark 12:** Input error MRAC has attracted much less attention in the adaptive control community than prediction-error MRAC. This, in spite of the fact that the former has the following significant advantages—already stressed in [52]: (i) As shown in (30), in contrast to input error MRAC, the regression model used in prediction error MRAC is bilinear. To overcome this difficulty it is necessary to overparameterize the identifier excluding the possibility of parameter convergence even when PE conditions are satisfied. (ii) The derivation of the error equation in prediction error MRAC fails if the input saturates, yielding erroneous updates in the identifier, see [17] for further discussion. These problems are conspicuous by their absence in input error MRAC.

**VII. NONLINEARLY PARAMETERIZED REGRESSION EQUATIONS**

In this section we provide an extension of the result of Proposition 2 to a class of NLPRE. We consider the case of CT separable NLPRE.

\[
y(t) = \phi^T(t)S(\theta),
\]

where \(S : \mathbb{R}^q \to \mathbb{R}^p\), with \(p \geq q\) is a known mapping of the vector of unknown parameters \(\theta\). Similarly to [9], [40], [55] the property that we exploit to achieve the estimation objective is monotonicity, which is defined with the following non-standard assumption.

**A.3** The mapping \(S(\theta)\) is strongly \(P\)-monotone in the sense that there exists a matrix \(P \in \mathbb{R}^{q \times p}\) such that

\[
(a - b)^T P [S(a) - S(b)] \geq \rho |a - b|^2 > 0, \quad \forall a, b \in \mathbb{R}^q,
\]

with \(a \neq b\) and for some \(\rho \in \mathbb{R}_{>0}\).

Notice that if \(p = q\) and \(P\) is positive definite Assumption A.3 reduces to the standard strong \(P\)-monotonicity of \(S(\theta)\) [12], [47].

**Proposition 7:** Consider the CT NLPRE (34) with \(S : \mathbb{R}^q \to \mathbb{R}^p\) satisfying Assumption A.3, that is (35) holds. Define the G+D interlaced estimator

\[
\hat{\theta}_g(t) = \gamma_g(t)\phi(t)[y(t) - \phi^T(t)\hat{\theta}_g(t)], \quad \hat{\theta}_g(0) = \theta_{g0} \in \mathbb{R}^p
\]

\[
\hat{\Phi}(t) = A(t)\Phi(t), \quad \Phi(0) = I_p
\]

\[
\hat{\theta}(t) = \gamma P\Delta(t)[Y(t) - \Delta(t)S(\hat{\theta}(t))], \quad \hat{\theta}(0) = \theta_0 \in \mathbb{R}^q,
\]

where \(\gamma_g(t) \in \mathbb{R}_{>0}\), \(\gamma \in \mathbb{R}_{>0}\), and we defined

\[
A(t) := -\gamma_g(t)\phi(t)\phi_t^\top(t)
\]

\[
D(t) := I_p - \Phi(t)
\]

\[
\Delta(t) := \text{det}(D(t))
\]

\[
Y(t) := \text{adj}(D(t))\{\hat{\theta}_g(t) - \Phi(t)\theta_{g0}\}.
\]

If \(\phi(t)\) is IE then \(\lim_{t \to \infty} \hat{\theta}(t) = 0, \quad (exp)\).

**Proof:** With some abuse of notation, define the signal

\[
\hat{S}(t) := \hat{\theta}_g(t) - S(\theta),
\]

whose derivative is given by

\[
\dot{\hat{S}}(t) = \gamma_g(t)\phi(t)[y(t) - \phi^T(t)\hat{\theta}_g(t)]
\]

\[
= -\gamma_g(t)\phi(t)\phi_t^\top(t)\dot{\hat{S}}(t)
\]

\[
= A(t)\hat{S}(t),
\]

where we used (36a), (34) and (37a) to get the first, second and third identities, respectively. Consequently, we get

\[
\hat{S}(t) = \Phi(t)\hat{S}(0),
\]

which may be rewritten as the NLPRE

\[
D(t)\hat{S}(t) = \hat{\theta}_g(t) - \Phi(t)\theta_{g0},
\]

\[4\]To avoid cluttering the notation, we restrict our presentation to the CT case, since as shown in [40] the extension to DT follows verbatim.
where we used (37b). Multiplying (38) by $\text{adj}\{D(t)\}$ we get the following NLPRE
\[
Y(t) = \Delta(t)S(\theta),
\]
where we used (37c) and (37d). Replacing (39) in (36c) we get
\[
\dot{\theta}(t) = -\gamma \Delta^2(t)P[S(\hat{\theta}(t)) - S(\theta)].
\]
To analyse its stability define the Lyapunov function candidate $U(\theta) := \frac{1}{2}|\theta|^2$, whose time derivative yields
\[
\dot{U}(t) = -\gamma \Delta^2(t)\dot{\theta}(t) - \theta^T P[S(\hat{\theta}(t)) - S(\theta)] \\
\leq -\gamma \Delta^2(t)\rho^2(\theta(t))^2 \\
= -2\rho^2 \Delta^2(t)U(t),
\]
where we invoked the Assumption A.3 of strong $P$-monotonicity of $S(\theta)$ to get the first bound. To complete the proof, we invoke the Comparison Lemma [22, Lemma 3.4] that yields the bound
\[
U(t + t_c) \leq e^{-2\rho^2 \int_0^{t_c} \Delta^2(s) ds} U(t),
\]
which ensures $\lim_{t \to \infty} \dot{U}(t) = 0$ (exp) if $\Delta(t)$ is PE. The latter condition follows from the assumption that $\phi(t)$ is IE and Lemma 4.

Remark 13: The estimation procedure of Proposition 7 differs from the one given in [40, Proposition 2] in several respects. First, while the latter uses the now classical DREM estimator procedure, in the former we propose to use the new G+D estimator of Proposition 2. The main advantage of this modification is that we replace the assumption of $\Delta(t) \notin \mathcal{L}_2$, imposed in [40], by the strictly weaker IE assumption of $\phi(t)$.

A second fundamental difference is that the monotonicity Assumption A.3 is imposed in the present paper to the original mapping $S(\theta)$, this differs with [40, Proposition 1] in two respects, first, a more general procedure to generate a new monotonic mapping, which involves a change of coordinates and a nonlinearity selection stage is propose in the latter. Second, the standard strict monotonicity assumption, instead of (35), is imposed to the new mapping.

VIII. SIMULATIONS

In this section we illustrate, via simulations, the main contributions of the paper.

A. Comparison of G+D and D+G estimators

To illustrate the result of Proposition 2, we consider in this subsection the problem of parameter estimation of a CT LTI system and choose, as an example, the system:
\[
y_p(t) = \frac{B(P)}{A(P)} u_p(t) = \frac{b_1 P + b_0}{P^2 + a_1 P + a_0} u_p(t),
\]
where $u_p(t) \in \mathbb{R}$ and $y_p(t) \in \mathbb{R}$ are the input and output signals, respectively. Following the standard identification procedure [52, Subsection 2.2] we derive the LRE (1) as follows
\[
y(t) := y_p(t), \quad \phi(t) := \left[ \frac{E(P)B(P)}{F(P)} \right] u_p(t),
\]
where
\[
F(P) := \frac{1}{R(P)} \begin{bmatrix} 1 \\ \vdots \\ R_{n-1} \\ b_0 \\ \vdots \\ b_{n-1} \end{bmatrix},
\]
with $R(P) = \sum_{i=0}^n r_i P^i, r_n = 1$, an arbitrary Hurwitz polynomial.

We compare the estimation of the parameters $\theta_i$ using the G+D interlaced estimator of Proposition 2 and the D+G scheme based on the generation of new LRE presented in [24], which is summarized below.

Consider the scalar LRE (1). Fix the constants $\lambda > 0$, $g > 0$, and define the signals
\[
\dot{Z}(t) = \mathcal{H}[\phi y](t), \quad \Psi(t) = \mathcal{H}[\phi \phi^T](t) \\
\dot{Y}(t) = \text{adj}\{\Psi(t)\} Z(t), \quad \Delta(t) = \det\{\Psi(t)\},
\]
where $\mathcal{H}[u] = \frac{n}{\pi} \int_{-\pi}^\pi |u(t)|$ is a linear filter. In [24] it is shown that if $\phi$ is IE then, $\Delta$ is IE and the $q$ scalar LREs
\[
\gamma_i(t) = \Delta(t) \theta_i, \quad i \in \bar{q},
\]
hold. Now, define the dynamic extension.\(^5\)
\[
\dot{Z}(t) = -k z(t) + k \bar{\Phi}_1(t) \Psi(t), \quad z(0) = 0 \\
\dot{\zeta}(t) = \bar{A}(t) \zeta(t) + \bar{b}(t), \quad \zeta(0) = \text{col}(0,0) \\
\bar{\Phi}(t) = \bar{A}(t) \bar{\Phi}(t), \quad \bar{\Phi}(0) = \text{col}(1,0),
\]
and
\[
\bar{A}(t) := \begin{bmatrix} 0 & -k \Delta(t) \bar{\Phi}_1(t) \\ k \Delta(t) \bar{\Phi}_1(t) & -\bar{V}(t) \end{bmatrix}, \\
\bar{b}(t) := \begin{bmatrix} -k \Delta(t) \bar{\Phi}_1(t) z(t) \\ [\bar{V}(t) - k] z(t) \end{bmatrix},
\]
where
\[
\bar{V}(\bar{\Phi}) := \frac{1}{2} (\bar{\Phi}_1^2 + \bar{\Phi}_2^2) - \beta,
\]
with $k > 0$ and $\beta > \frac{1}{2}$. Then, the new LRE
\[
\dot{Y}(t) = \bar{\Phi}_2(t) \theta,
\]
holds with
\[
\dot{Y}(t) := z(t) - \zeta_2(t).
\]
Moreover, $\bar{\Phi}_2(t)$ is PE and $z(t), \zeta(t), \bar{\Phi}(t)$ are bounded. Hence, using the standard gradient descent adaptation
\[
\dot{\theta}(t) = \kappa \Phi(t) \left( \dot{\Phi}(t) - \dot{\Phi}(t) \dot{\theta}(t) \right), \quad \kappa > 0,
\]
we get exponential parameter convergence. For further details of the D+G scheme see [24, Propositions 1 and 2].

\(^5\)To simplify the notation we omit the subindex $i$. 

To carry out the simulations, we use the system studied in [3, Section 5], that is $(a_0, a_1, b_0, b_1) = (2, 1, 1, 2)$ and choosing $r_1 = 20$ and $t_0 = 100$. This yields $\theta = \text{col}(98, 19, 1, 2)$ and for both estimators we propose an input signal that is not sufficiently rich, but generate a regressor $\phi(t)$ which is IE, namely
\[
u_p(t) = e^{-2t} + e^{-1.5t}.
\]

From Proposition 2 it is clear that the G+D estimator has only two tuning gains $\gamma$ and $\gamma_g$ that, as discussed in Remark 2, have a clear role in the transient behavior. From the material above we see that the D+G scheme has $\lambda, g, k, \beta$ and $\kappa$, whose impact on the transitory is rather obscure. Simulation experience has shown that tuning the gains of the D+G estimator is a hard task and a bad selection can have profoundly adverse effects on the behavior of the estimation. On the other hand, the tuning of the G+D estimator is relatively straightforward. To illustrate these facts, we present below some comparative simulations of both schemes. We focus on two important parameters of both schemes, which are, $\gamma$ and $\beta$, where both play a central role in the generation of the “exciting signals” in the regressors of the new LREs. Hence, we fixed the gains of the D+G scheme to $\lambda = g = 1, k = 0.4$ and $\kappa = 10$, while for the G+D estimator we use $\gamma = 200$ and propose the following different values for $\beta$ and $\gamma_g$
\[
\beta = \{0.65, 0.8, 1.1, 1.5\} \\
\gamma_g = \{2500, 2900, 3300, 3800\}.
\]

Using the same set of gains for all estimated parameters, all initial conditions $\hat{\theta}_i(0) = 0$ in both estimators and $\theta_{g_0} = \text{col}(0.4, 0.2, 0.5, 0.5)$ for the G+D one.

The simulation results, which corroborate the claims above, are shown in Figs. 1–4, where we depict the behavior of each $\hat{\theta}_i(t)$ for each value of $\beta$ and $\gamma_g$, distinguishing them by the color in the Figures. It is appreciated that the convergence of the G+D has a clear monotonic behavior with respect to $\gamma_g$. The convergence of the D+G one is also faster as $\beta$ increases but this introduces a “dead-time” in the response. We also see from Figs. 3 and 4 that if $\beta > 1$, the D+G scheme generates an unusual oscillatory behavior around zero of non-negligible amplitude—see the difference in the scales of the boxed regions. The reason for the appearance of both undesirable effects is not clear and does not follow from the theoretical analysis in [24].

**Remark 2**. We fixed the gains of the D+G scheme to $\lambda = 65, g = 2900, k = 0.5$ and for both estimators we propose an input signal that is not sufficiently rich, but generate a regressor $\phi(t)$ which is IE, namely
\[
u_p(t) = e^{-2t} + e^{-1.5t}.
\]

The simulation results, which corroborate the claims above, are shown in Figs. 1–4, where we depict the behavior of each $\hat{\theta}_i(t)$ for each value of $\beta$ and $\gamma_g$, distinguishing them by the color in the Figures. It is appreciated that the convergence of the G+D has a clear monotonic behavior with respect to $\gamma_g$. The convergence of the D+G one is also faster as $\beta$ increases but this introduces a “dead-time” in the response. We also see from Figs. 3 and 4 that if $\beta > 1$, the D+G scheme generates an unusual oscillatory behavior around zero of non-negligible amplitude—see the difference in the scales of the boxed regions. The reason for the appearance of both undesirable effects is not clear and does not follow from the theoretical analysis in [24].

**B. Application of G+D MRAC to Rohrs’ Examples**

In this subsection, we evaluate the performance of the G+D MRAC of Proposition 6. In particular, we consider the scenarios of [50], which are widely used as benchmarks to study the robustness of adaptive controllers vis-à-vis unmodeled dynamics and noise.

We consider a first-order plant
\[
y_p(t) = \frac{2}{p+1} u_p(t),
\]
and a reference model
\[
y_m(t) = \frac{3}{p+3} r(t).
\]

In this case the ideal controller gains are $\theta = \text{col}(1.5, -1)$. We adopt the reference signals proposed in [50], that is $r_1(t) = 0.3 + 18.5 \sin(16.1t)$ and $r_2(t) = 2$. According to [52], the reference $r_1(t)$ is “sufficiently rich” for a plant with two unknown parameters, thus the associated regressor of prediction error MRAC (31) satisfies the PE condition. On the other hand, for the reference $r_2(t)$, the PE condition is not satisfied and parameter convergence cannot be guaranteed.

First, we simulated the G+D MRAC for the ideal case in the absence of unmodeled dynamics. The initial conditions are set as $\hat{\theta}(0) = [0.1, 0.1]^T$ and all the others are selected as zero. The gains are chosen as $\gamma_g = 200$ and $\gamma = 100$, with the simulation results shown in Figs. 5 and 6. In both cases, we get satisfactory tracking performance, and the parameter estimation errors exponentially converge to zero even for the non sufficiently rich reference $r_2(t)$. 

**Fig. 1**: Transient behavior of $\hat{\theta}_1(t)$ and $\hat{\theta}_2(t)$ using the G+D scheme

**Fig. 2**: Transient behavior of $\hat{\theta}_3(t)$ and $\hat{\theta}_4(t)$ using the G+D scheme

**Fig. 3**: Transient behavior of $\hat{\theta}_1(t)$ and $\hat{\theta}_2(t)$ using the D+G scheme
As shown in [50]—see also [52, Subsection 5.2]—in the presence of unmodeled dynamics of the form
\[
y_p(t) = \frac{2}{P + 1} \cdot \frac{229}{P^2 + 30P + 229} |u_p(t)|,
\]
prediction error MRAC will diverge for both reference signals.

To assess the robust performance of G+D MRAC and verify the robustification claims of Proposition 4 we simulated the estimator, for the PE reference \( r_1(t) \), with a constant gain \( \gamma_g = 100 \), and a time-varying gain \( \gamma_g(t) = 100 + 0.1 \times t \), both with \( \gamma = 1000 \). From Fig. 7, we see that the constant gain adaptive controller is unstable. On the other hand, using a time-varying gain \( \gamma_g(t) \) guarantees signal boundedness of the closed-loop with very good parameter estimation as shown in Fig. 8. It should be underscored that, due to the presence of the unmodeled dynamics, parameter convergence to the ideal plant values is no guarantee for stability.

Finally, we considered the case of unknown sign of the high-frequency gain \( k_p \). To this end, we test the same controller with \( \gamma_g = 200 \) and \( \gamma = 100 \), but applying to the plant
\[
y_p(t) = \frac{-2}{P + 1} |u_p(t)|.
\]
The simulation results, given in Fig. 9, show that the proposed method still guarantees exponential convergence of both the state and estimation errors.

### C. Disturbance rejection with a NLPRE

In this section, simulations of the parameter estimation for the rejection of sinusoidal disturbances problem formulated in Section V-B are presented. First, we recall that using Proposition 5 we can generate from the scalar perturbed LREs...
We consider the following three different signals of the input signal on the performance of the estimator of $\theta$ used in all simulations are $\omega$ to estimate this unknown parameter.

We verify $A.3$, that is, $P\nabla S + (P\nabla S)^T = 2I_2$, so that, $S(\theta, \omega)$ is strongly $P$-monotone. Hence, it is possible to apply the estimation algorithm for NLPRE of Proposition 7.

Towards this end, we notice that this choice of $P$ allows us to consider as unknowns only the first and third elements of the mapping $S(\theta, \omega)$. For future reference, we rename them $\theta_1 := \theta$, $\theta_2 := \frac{1}{\omega^2}$.

We present now simulations of the estimator in Proposition 7 to estimate this unknown parameters.

We fix the disturbance $\xi = 0.5 \sin(\omega t + 0.4)$ with unknown frequency $\omega = 5$ and $\theta = 5$. The initial conditions that were used in all simulations are $\theta_0 = \text{col}(0.4, 0.2, 0.5)$ and $\theta_0 = \text{col}(0.2, 0.4)$. Besides, we set the tuning gains to $\gamma = 1$, $\gamma_0 = 0.9$ and $\gamma = 150$. To evaluate the effect of the richness content of the input signal on the performance of the estimator of Proposition 7 we consider the following three different signals $\Delta(t)$, namely:

$$\Delta_a(t) = \begin{cases} 1 & t \in [0, 4] \\ 0 & t > 4 \end{cases}$$

$$\Delta_b(t) = \frac{1}{0.2 + t^2}.$$  

$$\Delta_c(t) = 5 \exp^{-0.2t} \cos \left( \frac{\pi}{4} t \right).$$

Even though the three signals are not PE and belong to $L_2$, they are IE, thus the estimation (without overparametrization) of $\theta_1$ and $\theta_2$ is guaranteed using the estimator (36). Also, it is clear that the richness content of the signal above increases from the first to the last one. Fig. 10 corroborates this fact, where the line color distinguishes the signal $\Delta_a(t)$ that is used. On the other hand, Figs. 11 and 12 show the transient behavior of the signals $(\Phi_\xi)_{2,1}(t)$, $\Omega_{2,1}(t)$ and $\Omega_{2,2}(t)$ of the new LRE (27) and the signal $\Delta(t)$ of the interleaved estimator defined in (37c) in Proposition 7. The latter plot clearly confirms our claim regarding the richness of the signals $\Delta(t)$.

**IX. Conclusions and Future Research**

In this paper we have provided a solution to the problem of designing an on-line estimator that ensures GES of the PEE under the weakest assumption that the LRE is identifiable. Moreover, we have shown that, imposing a constraint on the adaptation gain of the first estimator, we prove that the scheme is robust to external disturbances and (not necessarily slow) parameter variations. We also proposed a variation of this estimator that rejects sinusoidal disturbances with unknown internal model and shown that the procedure is applicable to a well-defined class of NLPRE. Finally, we showed that, applying the proposed estimator in the MRAC problem, ensures relaxes the assumption of known sign of the high frequency gain.

Our current research efforts are oriented in the following directions.

- Application of the G+D estimator to the problem of state observation of state-affine systems as done in [43], from
which is clear that the conditions for convergence will be relaxed and the robustness properties improved. In particular, we are interested in the case of time-varying systems with unknown parameters as done in [8].

- Extend the material of Section VI to address other issues arising in standard MRAC. For instance robustness to unmodeled dynamics and the required prior knowledge for the multivariable case [15], [54]. The simulation results of Section VIII, that reveal some robustness of the new scheme with respect to the classical counterexamples of [50], being quite encouraging. However, a deeper understanding of the instability mechanisms, partially revealed in [5], is required.

- Extend the disturbance rejection result of Proposition 5 to the case of multiple frequencies. We have available a solution for two frequencies but the generalization to more frequencies is still to be worked out.

- Proceed with the comparative study of the proposed G+D estimator and the D+G one proposed in [24]. As shown in Section VIII the behavior of G+D is “monotonic” with respect to the tuning gains γ and γg. On the other hand, simulation evidence has shown that D+G has a more “erratic” dependence on λ, g, α, β, k and κ, hence the commissioning procedure of the former is “easier”. In any case, a better understanding of their similarities/differences is needed.

- In [57] a procedure to “mix” the estimates ˆθγ(t) and ˆθ(t) that preserves the main properties of Proposition 2 is proposed. The advantages of this new modification are yet to be clarified.

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In this appendix we prove that Φ in IE implies that the signal ∆ generated according to the construction of Proposition 2 is PE. The CT and DT cases are given in Lemmas 3 and 5, respectively.

Instrumental for the establishment of the CT claim is the following result from [13, Lemma 1].

**Lemma 2:** Let \( w(t) \in \mathbb{R}^q \) be the solution of

\[
\dot{w}(t) = -\gamma_g(t)\phi(t)^T w(t), \quad w(0) = w_0 \in \mathbb{R}^q,
\]

with \( \phi(t) \) being IE and \( \gamma_g(t) \in \mathbb{R}_{>0} \) being continuous and bounded. Define \( W(w) := \frac{1}{2} |w|^2 \). Then, it holds that

\[
W(0) - W(t_c) \geq \frac{\bar{\gamma}_g C_c}{1 + \bar{\gamma}_g^2 c^2 \phi_M^2} \gamma_g(t_c) |w_0|^2,
\]

where \( \Phi := \max_{t \in [0, t_c]} |\phi(t)|^2 \), \( \bar{\gamma}_g := \min_{t \in [0, t_c]} \gamma_g(t) \), and \( t_c, C_c \) are given in Definition 1.

**Lemma 3:** [CT case] Let \( \Phi(t) \) be the solution of (3b), (6) with \( \phi(t) \) bounded for \( t \in [0, t_c] \) and IE. Then, there exists \( \epsilon \in (0, 1] \) such that

\[
|\Delta(t)| \geq \epsilon^q
\]

for all \( t \geq t_c \). Consequently, \( \Delta(t) \) is PE.

**Proof:** Consider the LTV system (42) for an arbitrary nonzero \( w_0 \), whose fundamental matrix is \( \Phi(t) \). Hence, \( w(t) = \Phi(t)w_0 \) and

\[
W(t) = \frac{1}{2} |w(t)|^2 = \frac{1}{2} |w_0|^2 \Phi(t)^T \Phi(t)w_0
\]

and

\[
W(0) - W(t_c) = \frac{1}{2} |w_0|^2 \left[ I_n - \Phi(t_c) \Phi(t_c)^T \right] w_0.
\]

Invoking (43), we get the bound

\[
W(t_c) = \frac{1}{2} |\Phi(t_c)w_0|^2 \leq \frac{1}{2} \left( 1 - \frac{\bar{\gamma}_g C_c}{1 + \bar{\gamma}_g^2 c^2 \phi_M^2} \right) |w_0|^2.
\]

Since \( W(t) \) is non-increasing, it immediately follows that

\[
|\Phi(t)w_0|^2 \leq \left( 1 - \frac{\bar{\gamma}_g C_c}{1 + \bar{\gamma}_g^2 c^2 \phi_M^2} \right) |w_0|^2, \quad \forall t \geq t_c.
\]

Further, as \( w_0 \) is an arbitrary nonzero vector, it follows that the spectral radius of \( \Phi(t) \) satisfies

\[
\rho \{ \Phi(t) \} \leq \sqrt{1 - \frac{\bar{\gamma}_g C_c}{1 + \bar{\gamma}_g^2 c^2 \phi_M^2}}, \quad \forall t \geq t_c.
\]

where we used the fact that \( \phi_M t_c \geq C_c \) —which follows from Definition 1—and thus

\[
\rho \{ \Phi(t) \} \leq \frac{\gamma_c C_c}{1 + \bar{\gamma}_g^2 c^2 \phi_M^2} \in (0, 1).
\]

Consequently, each eigenvalue \( \lambda_i, i \in \bar{q} \) of \( I_n - \Phi(t) \) satisfies

\[
|\lambda_i(I_n - \Phi(t))| \geq 1 - \sqrt{1 - \frac{\bar{\gamma}_g C_c}{1 + \bar{\gamma}_g^2 c^2 \phi_M^2}}, \quad \forall t \geq t_c.
\]
The proof is completed recalling that $\Delta(t) = \det\{I_q - \Phi(t)\}$ and the fact that a scalar function whose lower bound converges to a non zero value is PE.

Instrumental for the establishment of the DT claim is the following result.

**Lemma 4.** Let $w(k) \in \mathbb{R}^q$ be the solution of

$$w(k + 1) = \left( I_q - \frac{\phi(k)\phi^T(k)}{\gamma_g(k) + |\phi(k)|^2} \right) w(k), \quad w(0) = w_0,$$

(45)

with $\phi(k)$ being IE and $\gamma_g(k) \in \mathbb{R}_{>0}$. Then, there exists $\alpha_0 \in (0, 1)$ such that

$$|w(k)| \leq \alpha_0 |w_0|, \quad \forall k \geq k_d$$

(46)

with $k_d$ given in Definition 1.

**Proof:** It is observed that for all $\bar k \geq k_d$,

$$\sum_{k=0}^{\bar k} \frac{\phi(k)\phi^T(k)}{\gamma_g(k) + |\phi(k)|^2} \geq \sum_{k=0}^{k_d} \frac{\phi(k)\phi^T(k)}{\gamma_g(k) + |\phi(k)|^2} \geq \cdots$$

$$\cdots \sum_{k=1}^{k_d} \frac{\phi(k)\phi^T(k)}{\gamma_g(k) + \phi_{M_k}} \geq \frac{C_d}{\gamma_g + \phi_{M_k}} I_q > 0,$$

where the second inequality is obtained by defining $\phi_{M_k} := \max_{k \in [0, k_d]} \|\phi(k)\|^2$ and the third is obtained by using the IE condition of $\phi(k)$ and defining $\gamma_g := \max_{k \in [0, k_d]} \gamma_g(k)$. Therefore, by recalling [53, Proposition 3.3], (46) can be concluded.

**Lemma 5:** [DT case] If $\phi(k)$ is IE, then there exists $\delta_d \in \mathbb{R}_{>0}$ such that

$$|\Delta(k)| \geq \delta_d, \quad \forall k \geq k_d.$$

Consequently, $\Delta(k)$ is PE.

**Proof:** Observe that

$$h^T D(k) h = h^T [I_q - \Phi(k)] h > 0, \quad \forall h \in \mathbb{R}^q \setminus \{0\} \Rightarrow |\Delta(k)| \in \mathbb{R}_{>0}$$

and the left-hand-side inequality holds if

$$|\Phi(k) h| < |h|, \quad \forall h \in \mathbb{R}^q \setminus \{0\}.$$ 

Whence, we will prove the lemma showing that, if $\phi(k)$ is IE, there holds

$$|\Phi(k) h| \leq \alpha_0 |h|, \quad \forall k \geq k_d, h \in \mathbb{R}^q$$

(47)

with $\alpha_0 \in (0, 1)$ given in (46). It is also noted that $\Phi(k)$ is the fundamental matrix of the system (45), which implies $w(k) = \Phi(k) h$ with $w_0 = h$. In this way, the proof reduces to show

$$|w(k)| \leq \alpha_0 |w_0|, \quad \forall k \geq k_d, w_0 \in \mathbb{R}^q,$$

which clearly is true by Lemma 4.

Therefore, there holds (47) and thus $|\Delta(k)| \geq (1 - \alpha_0)^q$ for all $k \geq k_d$, completing the proof.

---

**APPENDIX II**

**LIST OF ACRONYMS**

| Acronym | Description |
|---------|-------------|
| BIBO    | Bounded-input bounded-output |
| CT      | Continuous-time |
| DREM    | Dynamic regressor extension and mixing |
| DT      | Discrete-time |
| D+G     | DREM plus GPEBO |
| GPEBO   | Generalized parameter estimation based observer |
| GES     | Global exponential stability |
| G+D     | GPEBO plus DREM |
| IE      | Interval excitation |
| KP      | Key problem |
| LRE     | Linear regressor equation |
| LTI     | Linear time-invariant |
| LTV     | Linear time-variant |
| MRAC    | Model reference adaptive control |
| NLPRE   | Nonlinearly parameterized regressor equations |
| PE      | Persistent excitation |
| PEE     | Parameter error equations |