Calibrating Milky Way dust extinction using cosmological sources

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ABSTRACT

Aims. To constrain the light extinction properties of Milky Way dust.

Methods. We investigated the correlations between dust column density as inferred from infrared data and the observed colours of celestial objects at cosmological distances with low levels of colour dispersion.

Results. Results derived using colours of quasars, brightest central galaxies, and luminous red galaxies are broadly consistent, indicating a proportionality constant between the reddening $E(B-V) \equiv A_B - A_V$ is 0.02 with $A_V \sim 0.06$, see Fig. 1. In this paper, we re-investigate the relationship between the Milky Way reddening and the dust maps generated in Schlegel et al. (1998) using colours of quasars (QSOs) observed within the Sloan Digital Sky Survey (SDSS). We also compare with results from SDSS observations of brightest central galaxies (BCGs) and luminous red galaxies (LRGs). Compared with the original investigation in Schlegel et al. (1998), the improved statistics and multiple band photometry give us the possibility to decrease the errors on a linear relation while simultaneously constraining the properties of Milky Way dust, in this case the total to selective extinction ratio $R_V \equiv A_V/E(B-V)$.

We also performed a simple check to investigate if there are deviations from the linear relationship, in other words whether regions with high dust column density have more or less reddening than the linear prediction. No such effect is found, nor do we find any firm evidence of scatter in the dust properties as a function of the line-of-sight.

While noting that there is still room for non-linear behaviour, as well as scatter in dust properties at angular positions and scales not probed by the data, we conclude that at the positions and scales studied in this paper, it is a fair assumption that the reddening is linear to the dust column density $D^T$ (in units of MJy/sr) with proportionality constant $p = E(B-V)/D^T \sim 0.02$ and $R_V \sim 3$. The relative errors are of the order of 10% and are mainly limited by uncertainties in the filter response functions and the source spectral templates.

This agrees well with the results of Berry et al. (2011) where SDSS photometry of 73 million stars is used to obtain the dust extinction towards each star giving a three-dimensional dust map of the Milky Way in the direction of the observed stripes. The results from this analysis are consistent with $R_V = 3.0$, with a statistical and systematic error of 0.1 and a scatter of 0.2.

In Peek & Graves (2010), the colours of passively evolving galaxies from the SDSS are used to obtain corrections to the results of Schlegel et al. (1998). They find that the reddening in regions of low dust temperature is underpredicted with maximum deviations in $E(B-V)$ of 45 magnitudes.

Schlafly et al. (2010) use the blue edge of SDSS main-sequence stars to recalibrate the reddening induced by the dust maps of Schlegel et al. (1998). They find that the reddening is best described by a Fitzpatrick extinction curve with $R_V = 3.1$, whereas the extinction curve originally employed in Schlegel et al. (1998) generically overestimates the reddening. Similar results were obtained in Schlafly & Finkbeiner (2011) where very similar methods were used, but the reddening was measured by comparing the observed colours of SDSS stars to the colours predicted from their spectral features.

In the following, all limits are given as 95.4% confidence levels (CL) corresponding to $\chi^2 = \chi^2_{\text{min}} + 3.98$ for one parameter and $\chi^2 = \chi^2_{\text{min}} + 6.16$ for two parameters.

2. Method

We used the temperature corrected dust maps from Schlegel et al. (1998) to derive a reddening map by correlating the dust maps to the observed colours of celestial objects, preferably with a low intrinsic colour dispersion. The dust map, given as the dust column density $D^T$ in units of MJy/sr, is derived from Infrared Astronomical Satellite (IRAS) data, in principle allowing for a resolution of approximately five arcminutes. However, these data are calibrated and temperature corrected using degree resolution Cosmic Background Explorer (COBE) Diffuse Infrared Background Experiment (DIRBE) data. We note that in Kohyama et al. (2010), IRAS data are used to estimate the dust temperature in order to reach a resolution of five arcminutes also for the temperature corrected data of the
Cygnus region. However, no all sky map data have yet been produced using this technique. As a first attempt, we assumed that the observed colour, e.g. $E(B-V)$, is proportional to the dust column density $E(B-V) = pD^T + m$, where we set out to derive the proportionality constant $p$, and $m$ represents an arbitrary offset that is marginalised over (possibly caused by calibration errors giving the wrong mean colour) and the dust column density is given in units of MJy/st, making $p$ dimensionless. In Schlegel et al. (1998), a value of the reddening parameter $R_V = 3.1$ was assumed, yielding a value of the proportionality constant $p = 0.0184 \pm 0.0014$ for Mg2-calibrated BCGs.

Note that this can, in principle, be done for any set of filters $X$ and $Y$

$$E(X-Y) = k_{XY}D^T + m_{XY},$$

and that multiple colour data thus can be used to infer dust properties.

2.1. Fitting for $R_V$ and $p$

Since SDSS data allow for multiple colour combinations, we are also able to fit for the properties of the Milky Way dust and relax any prior assumptions on the reddening parameter $R_V$. The SDSS $[u,g,r,i,z]$ filters have central wavelengths at [3543, 4770, 6231, 7625, 9134] Å and we thus expect $g - r$ and/or $u - r$ to correspond the most closely to $B - V$.

For each source and filter, we calculate the transmitted flux using a redshifted source template. This is done for a grid of dust models in order to compute how each dust model will affect a given colour. Note that in doing this, we need to calculate the transmitted flux for each dust model separately since $k$-corrections will depend on the dust model. In practice, we have created tables that list the necessary $k$-corrections over a grid of redshifts $A_V$ and $R_V$ for each filter that we interpolate between. After calculating how a given dust model affects each source template in each filter, we can compare with the actual observed colours $E(X-Y)$ allowing us to constrain the parameters $p$ and $R_V$.

Normalising such that zero column density corresponds to zero reddening, i.e. $m = 0$ in Eq. [1] defining $f(\lambda) = A(\lambda)/\lambda$ and using

$$A_V = pR_V D^T,$$

we can write

$$A_1 - A_2 = \frac{\Delta m}{D^T} = pR_V (f_1 - f_2),$$

and

$$d\left(\frac{\Delta m}{D^T}\right) = dpR_V (f_1 - f_2) + pdR_V (f_1 - f_2) + prdR_V (f_1 - f_2),$$

showing how uncertainties in $D^T$, $\Delta m$, and the dust extinction curve $f(\lambda)$ translates into uncertainties in $p$ and $R_V$. In the following, we will make use of the Milky Way extinction curves $f(\lambda)$ as parametrised in Cardelli et al. (1989) (CCM) and Fitzpatrick (1999) (FTZ), plotted in Fig. 2 together with the (arbitrarily normalised) SDSS filter transmission curves. The extinction curve of O'Donnell (1994) gives results very close to those derived using the CCM parameterisation.

We now turn to the expected degeneracy between $R_V$ and $p$ and how this can be broken using multiple colour data. In the CCM-model, we have $f = a + b/R_V$ and $f_1 = f_2 = (a_1 - a_2) + (b_1 - b_2)/R_V \equiv \Delta a + \Delta b/R_V$. Setting $d\left(\frac{\Delta m}{D^T}\right) = 0$ in Eq. 5 and putting $d(\Delta a) = d(\Delta b) = 0$, i.e. no systematic error in the extinction curves, we obtain

$$dp = \frac{-dR_V}{R_V + \Delta b/\Delta a}.$$\hspace{1cm}(6)$$

This can be integrated to give

$$p = p_0 \frac{R_{V0} + \Delta b/\Delta a}{R_V + \Delta b/\Delta a},$$\hspace{1cm}(7)$$

describing the relation between different values of $R_V$ and $p$ giving the same extinction per unit column density $\Delta m/D^T$. The degeneracy lines for the SDSS standard colours $u - g$ (dark red), $g - r$ (red), $r - i$ (orange) and $i - z$ (yellow) are plotted in Fig. 3 displaying the power of multiple colour data to break the degeneracy between $p$ and $R_V$. We see that $p$ is anti-correlated with $R_V$, if the difference $(A_X - A_Y)/A_Y$ is decreasing more slowly than $1/R_V$ with $R_V$, which is the case for the three reddest SDSS standard colours. The slopes of these curves compare well with the confidence contours derived from data in the forthcoming sections.

3. Error analysis

We first note that constant filter calibration errors do not affect our results, since we are marginalising over any constant offset between the observed and theoretically calculated colours. However, we may get systematic effects if the spectral template is wrong and/or if the shape of the filter response function is incorrectly modeled since this will change the effective wavelength of the filters. We assume that this will induce fractional uncertainties in the effective wavelengths of [0.02, 0.01, 0.01, 0.01, 0.01] in $[u,g,r,i,z]$ filters (Doi et al. 2010). We note that this error will be a true systematic error since the effect will be correlated between all source redshifts and colours. For comparison, for our QSO template, effective wavelengths will vary by...
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Fig. 2. Dust extinction curves for different values of \( R_V \) shown together with the (arbitrary normalised) SDSS filter transmission curves. \textit{Top:} Extinction curve for the CCM parametrisation. \textit{Bottom:} Extinction curve for the Fitzpatrick parametrisation.

Fig. 3. Degeneracy between parameters \( p \) and \( R_V \) for SDSS colours \( u-g \) (dark red), \( g-r \) (red), \( r-i \) (orange) and \( i-z \) (yellow). The degeneracy lines correspond to a reddening equivalent to \( p = 0.02 \) and \( R_V = 3.0 \). It is evident that multiple colours allow the degeneracy to be broken.

\(~ 2 \sim 3\% \) when varying the source redshift and dust parameters. Schlegel et al. (1998) calculate their reddening estimates to have a dispersion of \( \sim 16\% \). Attributing a major part of this dispersion to a scatter in the dust column density, we assume a fractional uncertainty of 10\% in the dust map values \( D^f \). Since this uncertainty is treated as a statistical error, and we are systematics dominated, the exact value of this fractional error does not affect the end result. A possible problem not taken into account in this paper is that the dust map may be contaminated by the extragalactic background at low values (Yahata et al. 2007).

We incorporate uncertainties using a Monte Carlo-based approach as follows:

- We pick a random value for the shift in the effective wavelength for each filter and compute the corresponding colour offset. We also randomise the value for \( D^f \) at each sky position within the given uncertainties.
- We derive the likelihood over a grid of fitted parameters. In doing this, we use the statistical sources of error from observational uncertainties and the intrinsic colour dispersion of the sources.
- We repeat \( N \) times (we use \( N \sim 50 \)) to obtain \( N \) likelihood functions.
- We compute the mean of the \( N \) likelihood functions.

In this way we take full account of the different correlations between sources of error, and the resulting confidence contours correctly incorporate both statistical and systematic effects. As noted in Sect. 1, the dominating source of error is the possibility of shifts in the effective filter wavelengths, originating in uncertainties in the filter response functions and the source spectral templates.

4. Quasar analysis

It is well known that QSOs form a relatively homogeneous group of objects in terms of colours and spectral features. They also have large flux at short wavelengths which means that we can use blue colours, e.g. \( u-r \) even for high redshift sources, which allows us to effectively break the degeneracy between \( p \) and \( R_V \) as shown in Fig. 3. We use the fifth edition of the Sloan Digital Sky Survey Quasar Catalog (Schneider et al. 2007) consisting of 77 429 objects. The area covered by the catalogue is 5 740 deg\(^2\), see Fig. 4. QSOs have also been used successfully to constrain the amount and properties of dust in galaxies at cosmological distances ("Ostman et al. 2006, 2008) and for intergalactic dust ("Ostman & M"ortsell 2005).

We are primarily interested in the colours of QSOs, so it is important to understand how the objects in the SDSS QSO catalogue are chosen in order to assess whether or not there are selection effects that need to be considered. The QSO candidate selection is based on the position of sources in multidimensional SDSS colour space. In terms of colours, QSOs are mainly identified as being outside the four-dimensional stellar loci points in \([u,g,r,i,z]\) space (Richards et al. 2002). Since the method is purely differential, any general colour bias will not affect the end result; however, we need to make sure that the selection does not depend on the dust column. Since the selection is made after correcting for the extinction maps in Schlegel et al. (1998), assuming that these provide a fair model for the extinction, there is no reason that the selection should be severely affected by Milky Way dust extinction. By applying an extra colour cut of included objects where the (uncorrected) colour is less than 2 \( \sigma \) off the mean value, as evaluated in bins of different dust column
density, we are guaranteed a sample not correlated with the dust
column, unless initial dust corrections are grossly incorrect.
Only QSOs in the redshift range $0.5 < z < 3.0$ are considered
as QSOs outside this range are less accurately described by the
spectral template used, created by combining the Hubble Space
Telescope radio-quiet composite spectrum (Telfer et al. 2002)
with the SDSS median composite spectrum (Vanden Berk et al.
2001) as described in Östman et al. (2006, 2008). Generally, all
colours are relatively homogeneous at $z < 2$. Typically, at $z \sim 1$,
we have $\sigma(u-g) = 0.15$, $\sigma(g-r) = 0.13$, $\sigma(r-i) = 0.093$, and
$\sigma(i-z) = 0.11$. However, when the Lyman-$\alpha$ emission is being
redshifted into the observed filters, the dispersion increases. We
thus cut the $u$-band data at $z > 3000/1210 - 1 \sim 1.5$ and the
g-band data at $z > 4000/1210 - 1 \sim 2.3$.

In order to display the close to linear relationship between
the observed QSO colour and the dust density, in Fig. 5 we
show the observed QSO $g - r$ colour as a function of galac-
tic dust column density at redshifts $z \sim 0.95$, $z \sim 1.72$ and
$z \sim 2.48$. Results for the full fit of $p$ and $R_V$ are shown in
Fig. 6 for the CCM dust extinction parametrisation (top panel)
and the Fitzpatrick parametrisation (bottom panel). Solid colours
correspond to 95.4 \% CL for two parameters. Our constraints
are $p = 0.0205 \pm 0.0015$ and $R_V = 3.10 \pm 0.43$ (CCM) and
$p = 0.0190 \pm 0.0015$ and $R_V = 3.25 \pm 0.37$ (FTZ) (95.4 \% CL for
one parameter). Although consistent at the 95.4 \% CL, compared
to the $r-i$ and $i-z$ colours, we have on average less extinction in
$u-g$ and more extinction in $g-r$ for the CCM parametrisation.
For Fitzpatrick dust, compared to $g-r$ and $r-i$, $u-g$ has less
extinction and $i-z$ more extinction on average.

Albeit with very limited resolution, multiple colour data al-
low us to derive the dust extinction curve directly from the
data, without the need to assume any specific dust extinction
parametrisation. For each colour, we fit the slope $k_{XY}$ in Eq. 2
apply the necessary $k$-corrections, and normalise the extinction
to zero in the V-band. In Fig. 7 we show the dust law as derived from QSO colours compared to the extinction curves for some fiducial dust models. Assuming the errors to be accurately represented by the uncertainty of the slope in the linear fit, the error bars are too small to be visible. However, this assumption is valid only in the case that the relation between the reddening and the dust column density indeed is linear and constant over the field. Furthermore, the dominating source of error of possible systematic shifts in the effective central filter wavelengths have not been included. Note that since the normalisation is arbitrary, only the shape of the derived extinction curve is relevant.

5. Brightest central galaxies analysis

The brightest galaxies at the centre of galaxy clusters are empirically known to be relatively homogeneous in colour. Since high-luminosity early-type galaxies are generally redder than low-luminosity galaxies (usually attributable to higher metallicities), the colour dispersion can be further reduced using colour-magnitude relations. In this paper, however, we defer from employing any such calibration.

We have used data from the MaxBCG catalogue of 13 823 galaxy clusters from the Sloan Digital Sky Survey [Koester et al. 2007]. In order to extract the full photometric data for the brightest central galaxies (BCGs), we retrieved all objects within 6 arcseconds from the BCG positions. We managed to identify and retrieve \([u, g, r, i, z]\) data for 13 796 BCGs with redshifts \(0.046 < z < 0.37\). The angular positions of the BCGs are depicted in Fig. 8. Generally, all colours involving the u-band are not very well-determined. Typically (at \(z \sim 0.2\)), we have \(\sigma(u-g) = 0.32\), \(\sigma(g-r) = 0.059\), \(\sigma(r-i) = 0.029\), and \(\sigma(i-z) = 0.042\).

We cut all colours involving the the u-band at \(z > 0.15\) since the colour dispersion becomes very large at those redshifts. For consistency with the QSO and LRG analysis, we reject objects with uncorrected colours more than \(2\sigma\) off the mean value, as evaluated in bins of different dust column density. Results for fitting \(p\) and \(R_V\) to the BCG data are shown in Fig. 9. Our constraints are \(p = 0.0205 \pm 0.0030\) and \(R_V = 2.89 \pm 0.88\) (CCM), and \(p = 0.0177 \pm 0.0020\) and \(R_V = 3.42 \pm 0.48\) (FTZ) (95.4 \% CL for one parameter). Like the QSO constraints, compared to \(r-i\) and \(i-z\) that give \(p = 0.02\) for \(R_V = 3\), we have less extinction in \(u-g\) and more extinction in \(g-r\) for the CCM parametrisation. For Fitzpatrick dust, \(g-r\) and \(r-i\) give \(p = 0.02\) at \(R_V = 3\) while \(u-g\) has less extinction and \(i-z\) more extinction. In Fig. 10 we show the empirically derived dust law.

6. Luminous red galaxies

Since luminous red galaxies (LRGs) are also known to be homogeneous in terms of colours, we have extracted data for 114 151 LRGs from the SDSS data base. Generally, all colours involving the u-band are not well-determined. Typically, at \(z \sim 0.3\) we have \(\sigma(u-g) = 0.84\), \(\sigma(g-r) = 0.086\), \(\sigma(r-i) = 0.043\), and \(\sigma(i-z) = 0.056\). We therefore do not use any colours involving the u-band in this analysis. As discussed in Sect. 4 we reject objects with uncorrected colours more than \(2\sigma\) off the mean value, as evaluated at different dust column values, in order for the final sample to be unbiased by dust extinction. Fitting for
Fig. 8. **Top:** Angular positions of the SDSS BCGs used in this analysis with darker colours corresponding to higher observed surface densities of BCGs. **Bottom:** Rest-frame BCG template spectrum shown together with arbitrarily normalised SDSS filter transmission curves. Note that the same template is used for the LRG analysis.

$p$ and $R_V$ gives the results shown in Fig. 11. Our constraints are $p = 0.026 \pm 0.004$ and $R_V = 1.95 \pm 0.35$ (CCM), and $p = 0.019 \pm 0.0025$ and $R_V = 3.16 \pm 0.36$ (FTZ) (95.4% CL for one parameter). Compared to the QSO and BCG analysis, the absence of the $u$-band pushes best-fit values towards higher $p$ and lower $R_V$. Assuming $R_V = 3$, compared to $r - i$ and $i - z$ that give $p = 0.02$, we have more extinction in $g - r$ for the CCM parametrisation. For Fitzpatrick dust, $r - i$ gives $p = 0.02$ while $g - r$ and $i - z$ have more extinction. In Fig. 12 we show the derived dust law.

7. Discussion

As seen in the previous sections, there are slight differences in the preferred values of $p$ and $R_V$, depending on the dust parametrisation and the data set used. Generally, for the CCM parametrisation with $R_V = 3$, for $r - i$ and $i - z$ we get $p = 0.02$. We have more extinction (higher $p$) in $g - r$ and less in $u - g$. For Fitzpatrick dust, $r - i$ gives $p = 0.02$ while $g - r$ and $i - z$ have more extinction and $u - g$ less. How can we understand these differences? First, we need to understand the differences between the CCM and FTZ dust extinction parametrisations. In Fig. 13 we plot

$$\frac{f(\lambda)_{FTZ} - f(\lambda)_{CCM}}{f(\lambda)_{CCM}}.$$  

The reason that the we need higher $p$ for $i - z$ for the FTZ model is that the extinction curve is flatter there compared to the CCM model. We also see that it is steeper for $g - r$ which means that we get a lower $p$ for that colour. That is, we need a dust law that is slightly steeper for $g - r$ and flatter for $u - g$ compared to CCM. As for the differences between the data sets, comparing BCGs and LRGs, we see that the differences can be explained by having a flatter template spectrum for LRGs in the region covering $g - r$, compared to the BCG template. Comparing BCGs
and QSOs, the differences can be resolved by having either the BCG template flatter for \( u - g \) and \( g - r \) or the QSO template steeper in the same wavelength region.

8. \( R_V \)-variations

It is not unreasonable to assume that the distributions of grain sizes, and thus the resulting value of \( R_V \), is different in high- and low-density environments. Whether grains are generally larger or smaller in high-density environments depends on whether grain-grain collisions are more likely to cause coalescence or shattering. We therefore check for any evidence of correlations between \( R_V \) and the dust column density by dividing our sample into subsamples with different column densities. Note, however, that a high column density does not necessarily correspond to a high physical density. In Fig. 14 we show constraints from dividing the QSO sample into two regions with low \( (D^2 < 3 \text{ MJy sr}^{-1}) \) and high \( (D^2 > 3 \text{ MJy sr}^{-1}) \) dust column density. Our constraints for CCM are \( p = 0.020 \pm 0.0016 \) and \( R_V = 3.06 \pm 0.56 \) \( (D^2 < 3 \text{ MJy sr}^{-1}) \), and \( p = 0.019 \pm 0.0019 \) and \( R_V = 3.09 \pm 0.65 \) \( (D^2 > 3 \text{ MJy sr}^{-1}) \). For Fitzpatrick dust, \( p = 0.019 \pm 0.0016 \) and \( R_V = 3.28 \pm 0.37 \) \( (D^2 < 3 \text{ MJy sr}^{-1}) \), and \( p = 0.018 \pm 0.0018 \) and \( R_V = 3.23 \pm 0.41 \) \( (D^2 > 3 \text{ MJy sr}^{-1}) \) (95.4% CL for one parameter). This simple test does not provide evidence for any correlation between the dust column density and dust properties.

8.1. \( \sigma(R_V) \) and \( \sigma(A_V) \)

If dust has spatially varying properties, depending on the size of the regions we would expect the colour dispersion to be higher behind regions with high column density \( D^2 \). For filters \( X \) and \( Y \) we can write the colour \( E(X-Y) \) as a combination of the intrinsic colour \( E_i(X-Y) \) and the dust induced colour \( E_d(X-Y) \)

\[
E(X-Y) = E_i(X-Y) + E_d(X-Y),
\]

\[
\sigma(E(X-Y))^2 = \sigma(E_i(X-Y))^2 + \sigma(E_d(X-Y))^2,
\]

where

\[
E_d(X-Y) \equiv \left( \frac{A_X}{A_V} - \frac{A_Y}{A_V} \right) A_V \equiv f_{XY}(R_V) A_V,
\]

and

\[
\sigma[E_d(X-Y)]^2 = \left[ \frac{R_V \partial f_{XY}}{f_{XY} \partial R_V} \right]^2 + \left[ \frac{\sigma_{R_V}}{R_V} \right]^2 + \left[ \frac{\sigma_{A_V}}{A_V} \right]^2 E_d(X-Y)^2
\]

\[
\equiv g[X,Y,R_V,\sigma(R_V),\sigma(A_V)] E_d(X-Y)^2.
\]

In Fig. 15 we plot the variance \( \sigma[E(X-Y)]^2 \) as a function of \( E_d(X-Y)^2 \) for each standard colour of the QSO sample, expecting the variance at zero \( E_d(X-Y)^2 \) to give the intrinsic colour
variance \(\sigma[E_i(X-Y)]^2\) and the slope to be related to the scatter in \(R_V\) and \(A_V\) through the function \(q_i[X, Y, R_V, \sigma(R_V), \sigma(A_V)]\) (using the CCM dust extinction parametrisation). As can be seen, there are no obvious correlations between \(\sigma[E(X-Y)]^2\) and \(E_d(X-Y)^2\) (or \(\sigma[E(X-Y)]^2\) and the dust column density \(D^T\)). Assuming \(\sigma_{R_V}/R_V\) and \(\sigma_{A_V}/A_V\) to be constant, although the results depicted in Fig. 15 indicate that this assumption may be too simplistic, we can constrain their values from the fitted straight line slopes. Results from this analysis are shown in Fig. 16. The bluer colours show no increase in the colour variance with dust column density, whereas \(i-z\) show a slight increase in the variance indicating a possible dispersion of \(\sigma_{A_V}/A_V\) at the 30 % level. This said, there are a few points that should be remembered: First, this is under the assumption that the relation between the variance and the colour is indeed linear, and that the data points are not correlated. Second, results for \(i-z\) are not consistent with the bluer colours that show no signs of a scatter in the dust properties, although it is possible that the scatter only manifests itself at longer wavelengths. Third, there is a weak binning dependence when calculating the dispersion as a function of \(E_d(X-Y)^2\). To summarize, there is no evidence in the QSO data for any additional scatter in the observed colours in regions with high column density, which indicates that dust properties are relatively homogeneous averaged over the size of spatial regions probed in this paper, except perhaps for a hint of slightly increased scatter at the reddest optical wavelengths probed by the QSO data.

9. Summary

By investigating the correlations between dust column density as inferred from infrared data and the observed colours of celestial objects at cosmological distances with low levels of colour dispersion, we constrain the properties of Milky Way dust. We expect analyses of the reddening of extragalactic sources to complement analyses of Galactic sources very well since they will have different systematic uncertainties, and in principle may also probe dust at different distances. Results from QSOs, BCGs, and LRGs colours are broadly consistent, indicating a proportionality constant between the reddening \(E(B-V) \equiv A_B-A_V\) and the dust column density \(D^T\) (given in units of MJy/sr) of \(p = E(B-V)/D^T \sim 0.02\) and a reddening parameter \(R_V \equiv A_V/E(B-V) \sim 3\) with relative uncertainties of approximately 10 %. In spite of the increased statistics, the fractional error of \(p\) is similar to that of \cite{Schlegel1998}, the reason being that we simultaneously fit for \(R_V\) and that the uncertainty is dominated by the possibility of systematic shifts in the effective filter wavelengths.

We note that, in accordance with the analyses in \cite{Schlafly2010} and \cite{Schlafly2011}, the FTZ extinction curve may be argued to be preferred before the CCM parametrisation, since it is able to fit QSO, BCG, and LRG data with similar normalisations and reddening parameters. However, as argued in Sect. [7] such differences may also be traced to uncertainties in the spectral template functions.

The data does not provide evidence for any dependence of the dust properties as a function of dust column density, nor for any sizable scatter in the dust properties on the scales probed by the data, except for a hint of scatter \(\sigma_{A_V}/A_V \gtrsim 0.3\) in the extinction properties of dust at the longest wavelengths probed by the data (corresponding to \(i-z\)). We have also checked that within the uncertainties there is no evidence of any systematic difference between the dust extinction properties in the northern and southern hemisphere. Assuming the dust maps of \cite{Schlegel1998} to be correct, it is fair to assume that we can correct for Milky Way dust by applying a proportionality constant \(p = E(B-V)/D^T \sim 0.02\) with reddening parameter \(R_V \sim 3\) to the dust maps.

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Fig. 14. Constraints on $p$ and $R_V$ when dividing data into regions with low ($D^T < 3$ MJy sr$^{-1}$) and high ($D^T > 3$ MJy sr$^{-1}$) dust column density. There is no evidence for any correlation between the dust column density and dust properties. Upper left: Results for CCM dust for all regions with $D^T < 3$ MJy sr$^{-1}$ corresponding to 95.4% CL for one parameter of $p = 0.020 \pm 0.0016$ and $R_V = 3.06 \pm 0.56$. Upper right: Results for CCM dust at $D^T > 3$ MJy sr$^{-1}$ giving $p = 0.019 \pm 0.0019$ and $R_V = 3.09 \pm 0.65$. Lower left: Results for Fitzpatrick dust at $D^T < 3$ MJy sr$^{-1}$ giving $p = 0.019 \pm 0.0016$ and $R_V = 3.28 \pm 0.37$. Lower right: Results for Fitzpatrick dust at $D^T > 3$ MJy sr$^{-1}$ giving $p = 0.018 \pm 0.0018$ and $R_V = 3.23 \pm 0.41$. 
Fig. 15. The variance $\sigma^2[E(X-Y)]$ as a function of $E_d(X-Y)^2$ for each standard colour of the QSO sample. The total variance at zero $E_d(X-Y)^2$ is the intrinsic colour variance $\sigma^2[E_i(X-Y)]$ and the slope is related to the scatter in $R_{V}$ and $A_{V}$.
Fig. 16. Constraints on constant $\sigma_{R_V}/R_V$ and $\sigma_{A_V}/A_V$ derived from the linear fits depicted in Fig. 15. The coloured regions correspond to 95.4% CL (red) and 68.3% CL (yellow) for two parameters.
