Enhancing and comparing methods for the detection of fishing activity from Vessel Monitoring System data.

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Résumé

Vessel Monitoring System (VMS) data provide information about speed and position of fishing vessels. This opens the door to methods of estimating and mapping fishing effort with a high level of detail. To address this task, we propose a new method belonging to the class of hidden Markov models (HMM) that accounts for autocorrelation in time along the fishing events and offers a good trade-off between model complexity and computational efficiency. We carry out an objective comparison between this method and two competing approaches on a set of VMS data from Denmark for which the true activity is known from onboard sensors. The DMKMG approach proposed outperformed the competitors approach with 6% and 15% more accurate estimates in the vessel-by-vessel and trip-by-trip case, respectively. In addition, these better performances are not paid in terms of computation time. We also showcase our method on an extensive dataset from Sweden. A quick (real-time) data processing has the potential to change how fisheries can be better harmonized to other utilisation of the seas and fill the gap between the local scales at which fishing pressure and stock depletion occurs with the large temporal and spatial scales of traditional fisheries assessment and management. The computer code developed in this work is made publicly available as an R package from http://www2.imm.dtu.dk/~gigu/HMM-VMS.

Keywords: Vessel Monitoring System (VMS), Model-based clustering; Hidden Markov Models; EM algorithm; depmixS4; fishing effort; Fisheries Management.

1 Background

The ecosystem approach to fisheries (EAF) increasingly requires spatially resolved fisheries data, and suggests the inadequacy of traditionally aggregated landing and effort data to characterize the impact of the fisheries on species, habitats and ecosystems (Jennings, 2005; Douvere and Ehler, 2009). Marine spatial planning, which is recognised as an essential step towards an
ecosystem-based management of the sea, demands for detailed knowledge on the spatial and
temporal distribution of human activities including fishing (Douvere, 2008).

The implementation of a system for satellite-based geographic monitoring of fishing vessels
(Vessel Monitoring System VMS), originally for control purposes, has made possible to track
and map fishing effort on a much finer scale. Over the past few years VMS data has become
more widely available for scientific purposes, and it is now part of the indicators listed in the
EU data collection framework (European Commission 2008). This enhanced the development
of new approaches and tools (e.g. Hintzen et al. 2012) to analyse this type of data, and the
establishment of international cooperation to facilitate regional and multinational analysis of the
distribution and intensity of the total fishing effort (ICES Working Group on Spatial Fisheries
Data 2014; Raetz and Holmes 2013). Most of the methodological work on VMS data has
focused on developing methods for dealing with inherent limitations of this data source (Lee et
al. 2010), given the original purpose of control when the system was enforced in the European
Union in January 2000. Main limitations include a partial coverage of the fishing activities
as VMS was initially adopted on vessels longer than 24 meters and progressively extended to
smaller vessels to comprise all units larger than 12 meters since 2012. From 1 January 2005
the recording frequency must be lower than two hours but it is rarely lower than one hour.
Moreover, VMS returns vessel positions regardless the activity performed by the vessel at that
particular time which may comprise steaming to and across fishing grounds, searching for the
catch, actual fishing and handling the catch.

Identification of behavioural states and inference on the activity of individual vessels from
discrete and semi-regular observations is a challenging issue. This is particularly true considering
that spatio-temporal dynamics of fishing effort are the result of the development of fishing tactics
and strategies of individual fishers which are highly context- and fishery-dependent (Salas and
Gaertner 2004). Fishers show an adaptive response to changes in resource abundance and
distribution, environmental conditions and market or regulatory constraints (Bastardie et al.,
2015) which may be potentially detected by changes in the VMS patterns. Given the impact
that misclassification of vessel activity may have on the estimation of fishing effort, the need for
rigorous and reliable methods for detection of fishing activity has attracted a large amount of
methodological research on VMS data (Lee et al. 2010; Vermard et al. 2010). Recent advances
in the statistical analysis of movement and behavioural data have opened opportunities for the
development of generic and more rigorous statistical approaches to the analysis of VMS data.
Improvements in distinguishing between different activities or states from VMS data would have multiple benefits. It would improve estimation of the effective fishing effort which can then be translated into a measure of fishing pressure on species and a measure of impact on habitats. Moreover, the different activities that characterise a fishing trip have usually different associated costs and revenues, and the amount of time spent in each has to be identified for an economic evaluation (Pelletier et al., 2009; Bastardie et al., 2015).

The common rationale behind methods of detecting fishing activity from Vessel Monitoring System (VMS) data lies in the fact that fishing activity results typically in slower and more erratic trajectories than steaming activity. This is clearly visible on the vessel trip shown on Fig. 1 and for which brief steaming periods at high speed between the harbor and a trawling area flank a long period of erratic movements at low speed.

A simple method to estimate the state of a vessel consists in setting a lower and an upper threshold on the speed modulus and estimate any ping or time step as fishing if it is within these thresholds. This method is implemented for example in the VMStools program (Hintzen et al., 2012) and became a widely used approach across Europe for processing VMS and logbook fisheries data (see also Lee et al., 2010 for review and further references). A potential weakness of this method is that disregards any auto-correlation in times, in particular, the facts that a vessel does not alternate between fishing and steaming states too frequently. In principle, this simple threshold method can be improved in many ways. Mills et al. (2007) proposed to combine both the speed modulus and directionality in such a threshold-based method. Bastardie et al. (2010) proposed to account for auto-correlation in time of vessel states by using a hidden Markov model (HMM). Peel and Good (2011) enhanced the previous method by accounting for other states than fishing and steaming such as exiting, entering and staying at an anchorage. They also take into account the time of day in the estimation. Gloaguen et al. (2014) proposed to model auto-correlation in time of the vessel speed within each state with an HMM model. This contrasts with the methods of Bastardie et al. (2010) that assumes that speeds within a fishing segment or within a steaming segment are independently distributed. Both Peel and Good (2011) and Gloaguen et al. (2014) harness parameter estimation to maximum likelihood computation and the EM algorithm. We note also that earlier works differ in terms of type of validation carried out. Mills et al. (2007) and Bastardie et al. (2010) confronting their estimates
to the known activity available from on-board camera or observers data, while \cite{Peel2011} and \cite{Gloaguen2014} restricting their investigation to the use of simulated data.

The premise of the present paper is that model-based clustering methods and in particular hidden Markov models offer a flexible yet computationally efficient framework for the estimation of vessel activity and deserve further attention. The work is developed in three steps. First, we investigate in detail the application of a new HMM model and several variants in their practical implementation. Second, we compared the new approach to a number of existing methods, including the threshold-based method implemented in VMStools and to the model-based method of \cite{Gloaguen2014}. To do so, we evaluate the accuracy of methods with the aid of on-board sensor validation data. Third, we implement the model found to be the most accurate on an extensive dataset from the Swedish bottom trawl fisheries operating in the Baltic and the Kattegat-Skagerrak. The Swedish bottom trawlers target a number of key demersal species in these areas and contribute to more than 20% of the total EU catch of cod (\textit{Gadus morhua}) in the Baltic and more than 25% of Norway lobster (\textit{Nephrops norvegicus}) catch in the Kattegat-Skagerrak area.

2 Material and methods

2.1 Data

To design our model and showcase its main features, we study here a set of VMS data consisting of 6430 fishing trips carried out in 2009 by 131 Swedish bottom trawlers operating in the Baltic (ICES subdivision 24-25) and the Kattegat-Skagerrak (ICES subdivision 20-21). VMS data are available from the Swedish Agency for Marine and Water Management with link to fishermen logbooks based on individual vessel signal and date-time information on the trip. According to information from the logbooks, VMS data associated to specific fleet segments (metier unit at the DCF level 5) were selected for particular combinations of gear type (i.e., otter-trawlers) and target species assemblage (i.e., demersal fish and crustaceans) as calculated from the reported landings.

In order to assess the accuracy of the methods studied here, we used a validation dataset of ten trips made by six Danish vessels and for which the actual state (fishing or steaming) is exactly known from sensors attached to the fishing gear \cite{Bastardie2010}. Both VMS datasets feature approximately a one-hour ping frequency.
2.2 Model proposed for the clustering of VMS data

We denote by \( v_t \) the linear speed (speed modulus in the direction of movement) and \( \omega_t \) the angular speed (change of direction per time unit) of the vessel at time \( t \). These variables can be either available directly from on-board sensors as part of the VMS monitoring system or derived numerically as finite differences from GPS positions available readily as VMS data. Following a standard approach in model-based clustering (McLachlan and Peel, 2000), we introduce a discrete variable \( s_t \) defined at each time step and encoding the unknown state of the vessel. In the simplest instance, \( s_t \) has two states coded as 1/2, standing for fishing and steaming, and we assume that conditionally on \( s_t = k \) (\( k=1 \) or 2), the vector \( (v_t, \omega_t) \) arises from a bivariate Gaussian distribution with state-specific mean \( \mu_k \) and variance matrix \( \Sigma_k \). The difference between \( \mu_1 \) and \( \mu_2 \) reflects the difference of sailing behavior during fishing and steaming. Because the data we analyze are not perfectly fitted by a two-cluster Gaussian mixture (cf. Fig. 2, left panel), we consider a class of models in which the number of states \( K \) can be larger than 2. For example, in a model where \( K = 3 \), a Gaussian component will typically fit fishing states while the two other components will fit the various steaming states, one for the high-speed steaming states and the other one for the various low-speed steaming states (cf. Fig. 2, right panel).

As pointed out earlier and illustrated by figure 1, a vessel does not alternate constantly between fishing and steaming states. Using a statistical phrasing, the sequence of states \( (s_1, ..., s_T) \) is not properly model by a sequence of independent variables. Injecting this information in the model is presumably a good way to increase accuracy of inference. We do this by mean of a Markov chain in which the probability of the next state depends on the current state. The set of conditional probabilities \( p_{i,j} = p(s_{t+1} = j | s_t = i) \) forms a matrix known as transition matrix (Ross and Pekez, 2007) and is part of the unknown quantities to be estimated. The model formed by the Gaussian components together with the Markov chain for the un-observed states belongs to the class of hidden Markov models (HMM) which is an important tool in science and engineering (Zucchini and MacDonald, 2009) to model heterogenous data, in particular multi-modal data. The particular models we consider here can be referred to as dependent mixture of \( K \) multivariate Gaussian models (DMKMG). To estimate parameters, we use likelihood maximization for HMMs by the EM algorithm implemented in the R package depmixS4 (Visser and Speekenbrink, 2010).

In a model with \( K = 2 \) components, the component estimated with the lowest mean speed
has to be interpreted as the component modeling Fishing states (cf. Fig. 2 for an illustration). As soon as we make use of a model that includes more than two components, the interpretation of the various components is not straightforward. A possible approach consists in imposing some constraints on the parameters at the estimation stage, for example constraining the mean speed of a certain component (that would stand for Fishing states) to be e.g. less than 3 knots. Pilot experiments along this line led to convergence issues with the EM algorithm. Besides, constraining parameters does not fully address the issue of the interpretation of the components. Therefore we did not pursue this approach. We propose instead an alternative strategy in which the parameters are not constrained during the parameter estimation stage and the labelling is performed after parameter estimation. It is based on the observation that the low-speed component for a 2-component model has a variance that is too large because it includes many Steaming states. Our labelling algorithm goes as follows: we label as Fishing the combination of components that has a joint empirical variance smaller than that of the low-speed states estimated in a 2-component model. If there is more than one combination of components that reduces the variance, we choose the one with a joint empirical mean closest to that of the low-speed component estimated in a 2-component model.

2.3 Competing methods considered here

We compare outputs of the method outlined above to those obtained first with a simple threshold-based method and implemented in the program VMStools [Hintzen et al., 2012] and second with the method proposed by Gloaguen et al. (2014). In the model of Gloaguen et al. (2014), the speed is parameterized in terms of a persistence speed (along the direction at the current time step) and a rotational speed (along the direction at the next time step). This model assumes the existence of two states (Fishing, Steaming) whose distribution is given by a Markov chain. Because it is reasonable to assume that within a state, the speed varies smoothly, Gloaguen et al. (2014) place a time dependence structure on each component of the speed vector in form of a first-order auto-regressive Gaussian process. Their model appears therefore as a time-dependent mixture of auto-regressive processes (DMARP). They carry out inference with the EM algorithm. In the sequel, we use the R code developed by these authors (M. Gloaguen, personal communication). A comparison to the model proposed by Peel and Good (2011) would have been relevant here but the code developed by these authors has not been made available.
2.4 Validation against fine resolved data

The model we outlined in section 2.2 can be implemented with any number $K$ of Gaussian components. It can be run either trip-by-trip (therefore with trip-specific parameters), or vessel-by-vessel, i.e. treating all trips of a vessel as a single long trip (therefore with vessel-specific parameters) or even all data as one trip, i.e. treating all the data as a single long trip (therefore with a common set of parameters for the whole dataset). Besides, the model can be implemented using both linear and angular speed or linear speed only (assuming the use of angular speed only would not make sense). This defines up to $K \times 3 \times 2$ sub-models. The first goal of the present analysis is to assess which sub-model yields the most accurate results. To do so, we run each sub-model on the validation dataset, obtain a vector of estimated fishing states and compare it to the true state available from on-board camera data. The congruence between the vector of true and estimated states is assessed numerically through four statistics:

\[
\text{Global match} = \frac{\# \{ \text{Estimated state} == \text{True state} \}}{\# \{ \text{Time steps} \}}
\] (1)

\[
F \text{ as } S = \frac{\# \{ \text{Fishing state estimated as Steaming state} \}}{\# \{ \text{Time step} \}}
\] (2)

\[
S \text{ as } F = \frac{\# \{ \text{Steaming state estimated as Fishing state} \}}{\# \{ \text{Time step} \}}
\] (3)

\[
\text{Unestimated state} = \frac{\# \{ \text{Missing estimated state} \}}{\# \{ \text{Time step} \}}
\] (4)

In equation [4], missingness corresponds to states for which inference failed at least partly, typically for short trips that bring up identifiability issues. In a first set of pilot computations, we compared accuracies obtained with both linear and angular speeds to those obtained with linear speeds only. We carried out comparisons trip-by-trip, vessel-by-vessel and all data together for $K = 2$. In a second set of analyses, we estimated vessels states with the DMKMG model using linear speeds alone with a number of Gaussian components $K$ ranging from 2 to 10. Again we carried out comparisons trip-by-trip, vessel-by-vessel and all data together. In a third set of analyses, we compared the accuracy of the sub-model we found to be the most accurate in the analysis described above, to that obtained with VMStools and the DMARP model [Gloaguen].
et al., 2014) described in the previous section.

2.5 Application to Swedish data

We use the model found to be the most accurate on the Danish validation data to map the fishing effort from the Swedish demersal trawl fleet in 2009. In a vector of estimated states \( \hat{s}_1, \ldots, \hat{s}_T \) we consider that a stretch of consecutive states estimated as Fishing states form an estimate of a trawling event. Once the fishing states are separated from the other states, we map the fishing effort for this period over the western Baltic Sea and the Kattegat-Skagerrak.

3 Results

3.1 Danish data

Under the DMKMG model, we observed that making use of angular speeds together with linear speeds does not bring any improvement, or even lead to lower accuracies, than estimates using linear speed only, all other options being equal. Using linear speeds only, the DMKMG model yielding the highest accuracy in terms of global match is a model with \( K = 3 \) Gaussian components. Figure 2 illustrates the fit of this model and the empirical and theoretical distributions of known steaming and fishing states from the validation dataset. Results show that an extra Gaussian component helps capturing some of the low speed steaming events. Note that the overlap of red and blue dots on the x-axis in the right panel is due to the use of a HMM that does not classify states according to the estimated density value (cf. Zucchini and MacDonald, 2009).

![Figure 2 about here.]

Accuracies of this best model together with those obtained with the competing methods are reported in table 1. In the sequel, unless specified otherwise, all results refer to this best model.

The DMKMG model (three Gaussian components, linear speed only) shows the highest level of global match followed by the threshold method implemented by VMStools and then by the DMARP model. In the 'all data together' comparison, both DMKMG and VMStools are able to estimate all the states and achieve a correct identification score of 81.5% and 79.0%, respectively. DMARP achieves only a 68.5% global match due to 1.5% of unestimated states and mostly because of approximately 30% of steaming as fishing mis-classifications compared to the 17.5% of
DMKMG and 19.6% of VMStools. In the vessel-by-vessel and trip-by-trip comparison DMKMG maintains same high levels of global match (i.e., 78.6% and 79.9%, respectively). On the contrary, VMStools shows a decrease in performances (i.e., 72.9% and 64.5%, respectively), comparable to the matching level achieved by DMARP, due to an increased number of unestimated states.

[Table 1 about here.]
3.2 Swedish data

Figure 3 represents the spatial distribution of the fishing effort of the Swedish bottom trawl fisheries in the western Baltic (ICES subdivision 24-25) and the Kattegat-Skagerrak area (ICES subdivision 20-21) in 2009. In the western Baltic, the Swedish fisheries is highly concentrated in the waters northern Bornholm Island between the subdivisions 24 and 25, and along the 50-70 m isobaths that stretch in the northern and eastern side of the Bornholm Basin. The central and deepest part of the Bornholm Basin is mostly free from bottom trawling effort. Fishing effort is widely distributed along the Swedish Kattegat and Skagerrak on soft bottoms in the range of 50-200 m depth. Small and medium size un-trawled patches are visible throughout the area. In the Skagerrak the fishing effort of the bottom trawl fisheries extends westward into the North Sea, bounding the southern side of the Skagerrak deep (>300 m depth) which is left free from bottom trawling.

[Figure 3 about here.]
4 Summary and discussion

We have implemented various methods to detect fishing activity from VMS data. The method that performs best on the Danish data is the dependent mixture with three components. Presumably because it achieves the best trade-off between model parsimony and complexity. On the Danish trawlers data, the classification error can be brought below 20% and the proportion of Fishing states estimated as Steaming around 1%. We note that these numbers are likely lower on the Swedish data as the vessels are similar but parameters estimation is performed on a much larger dataset. In supplementary material, we report further investigation about the poor performances obtained when using information about the change of direction of the vessel. It is seen that in the various parametrisations of the speed vector, information carried by directional data can not be used efficiently by Gaussian mixture models.

One of the question that triggered this work was whether logbook data could be used to increase accuracy of algorithm for the detection of vessel states. Here we observed that errors of methods based on VMS data only correspond to isolated low-speed steaming states. It seems therefore unlikely that information about trawling fishing activity present in the logbook data could prevent this type of errors.

Peel and Good (2011) suggest, and this is confirmed by a preliminary analysis of our validation data, that fishing events tend to occur preferentially at a specific time of the day (day time in our Danish data). It would therefore be natural to inject this property in our model by linking the vector of transition probabilities $(p_{i,j})_{j=1,...,K}$ to the time of the day $h_t$ using a multinomial logistic model (Visser and Speekenbrink 2010). However, the time of the day is a circular variable ($h_t = 0:00$ is equivalent to $h_t = 24:00$) and this is not well handled by the link functions implemented in depmixS4. This seems to be a promising direction to further improve accuracy of fishing detection methods in this context.

The computer code we used for the K-component Gaussian mixture model is based on the R package depmixS4 programmed in C. This allows us to obtain computing time of the order of ten minutes to carry out inference on the whole Swedish dataset (131 vessels, 6430 trips) on a standard single-processor computer. This opens the door to accurate real-time information about accumulated fishing effort for decision makers (Knaak et al. 2012 Needle and Catarino 2011). The increasing availability and coverage of high-resolution real-time data on vessels po-
sitions (i.e., VMS and Automatic Identification System) and catches (i.e., electronic logbook) is profoundly changing the way we look at fishing pressure and impact on the marine systems. This has the potential to change how fisheries can be managed in the future and better harmonized to other utilisation of the seas (the maritime spatial planning is central in the development of the MSFD as stated in the EU MSP 2014/89/EU directive). One of the main cutting edge use of such high resolution data consists in the potential reconciliation of the existing gap between the small and local scales at which fishing pressure and stock depletion may occur and the large temporal and spatial scales of traditional fisheries assessment and management (Bartolino et al., 2012). Recently, new ideas for real-time management of fisheries have been proposed (Holmes et al., 2011; Needle and Catarino, 2011). For instance, the real-time incentive (RTI) system is stimulating a discussion on alternative forms of fisheries management which may raise from application of these new technologies (Kraak et al., 2012). Among the different aspects at discussion there are issues concerning the potential response of fisheries and occurrence of unpredicted and unlikely behaviour in fishermen, uncertainty in the assessment of the resources and value of habitat features impacted by the fisheries, technical requirements involved in the monitoring and efficient real-time processing of extensive information on the fisheries behaviour and catch (Kraak et al., 2014). The DMKMG approach proposed showed a marginal improvement compared to the other methods in the 'all data together' conditions investigated, but it outperformed the competitors with 6% and 15% more accurate estimates in the vessel-by-vessel and trip-by-trip case, respectively. In addition, these better performances are not paid in terms of computation time which as the EM resorting to the depmixS4 package reduce computing time by one or two order of magnitude, moving forward the current limits of real-time computation of VMS for fisheries management.

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Table 1 – Accuracy in estimating vessels activity on Danish data. All values are percentages. Number in parentheses are obtained when counting unestimated states as errors. See text for detail.

| Method          | Error statistics | all data together | vessel-by-vessel | trip-by-trip |
|-----------------|------------------|--------------------|------------------|--------------|
|                 | Global match     | 78.99              | 77.61 (72.88)    | 78.48 (64.48) |
|                 | F as S           | 1.34               | 1.59 (1.49)      | 1.62 (1.34)  |
|                 | S as F           | 19.67              | 20.79 (19.52)    | 19.89 (16.39) |
|                 | Unestimated states| 0                  | 6.11             | 17.58        |

|          | Global match     | 69.6 (68.55)       | 70 (69.00)       | 66.57 (65.57) |
|          | F as S           | 0.3 (0.3)          | 3.47 (3.43)      | 2.11 (2.09)   |
|          | S as F           | 30.10 (29.66)      | 26.48 (26.08)    | 31.32 (30.85) |
|          | Unestimated states| 1.51               | 1.51             | 1.51          |

|          | Global match     | 81.52 (81.52)      | 78.87 (78.62)    | 80.25 (79.94) |
|          | F as S           | 0.94 (0.94)        | 1.30 (1.30)      | 2.10 (2.09)   |
|          | S as F           | 17.54 (17.54)      | 19.82 (19.76)    | 17.65 (17.58) |
|          | Unestimated states| 0                  | 0.32             | 0.39          |