Intermediate redshift calibration of Gamma-ray Bursts and cosmic constraints in non-flat cosmology

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ABSTRACT
We propose how to calibrate long gamma-ray burst (GRB) correlations employing intermediate redshift data sets, instead of limiting to $z \approx 0$ catalogs. To do so, we examine the most updated observational Hubble data (OHD) and baryonic acoustic oscillations (BAO). We exploit the model-independent technique of Bézier polynomial interpolation, alleviating de facto the well-known circularity problem affecting GRB correlations. To get constraints on cosmic parameters, using Markov chain Monte Carlo Metropolis algorithm, we distinguish cosmic parameters, using Markov chain Monte Carlo Metropolis algorithm, we distinguish

1 INTRODUCTION
GRBs are often challenged as possible distance indicators and their use in cosmology is currently debated (Luongo & Muccino 2021a). Their possible use is essential to highlight possible departures from the concordance $\Lambda$CDM paradigm, whose overall dynamics is described by six free parameters (Perivolaropoulos & Skara 2021), albeit at late times it can be well-approximated only by matter, $\Omega_m$, accounting for about the 30% of the total energy budget (Ratra & Peebles 1988). The cosmological constant density drives the universe to accelerate and its experimentally and statistically agreement is robust, albeit recent tensions have been raised (Di Valentino et al. 2021), indicating mild discrepancies from the standard model predictions1 (Sotiriou & Faraoni 2010). Moreover, in the concordance paradigm, one conventionally assumes a perfectly spatially-flat universe, supported by several observations2. In this sense, bounding spatial curvature even remains as an additional open caveat of cosmology (Ooba et al. 2018).

Intermediate and high redshifts data exceeding the redshift limits of supernovae Ia (SNe Ia) detectability, placed at $z \approx 2.3$ (Rodney et al. 2015), and of other cosmic indicators, in general at $z \gtrsim 3$, turn out to be essential in order to shed light into the nature of those constituents pushing up the acceleration of the universe3 and to disclose whether the $\Lambda$CDM model may be seen as the final scenario describing large-scale dynamics or a limiting case of a more general landscape (Capozziello et al. 2019, 2020). In this respect, GRBs could represent a plausible new class of cosmological indicators. These explosions are detectable up to $z = 9.4$ (Salvaterra et al. 2009; Tanvir et al. 2009; Cucchiara et al. 2011) and so attempts toward their use in cosmology as genuine cosmic indicators are currently highly debated. Essentially a way out to relate GRB photometric and spectroscopic properties is the missing puzzle piece (Amati et al. 2002; Ghirlanda et al. 2004; Amati et al. 2008; Schaefer 2007; Capozziello & Izzo 2008; Dainotti et al. 2008; Bernardini et al. 2012; Amati & Della Valle 2013; Wei et al. 2014; Izzo et al. 2015; Demianski et al. 2017a,b). Moreover,

1 For different perspectives about extensions of the standard model, see e.g. (Luongo & Muccino 2018; D’Agostino et al. 2022; Bellfiglio et al. 2022).
2 Cosmic microwave background observations, for instance, seem to favor this fact, having $\Omega_k = -0.001 \pm 0.002$ at 68% confidence level. Further, inflationary paradigms also require severe limits on spatial curvature (Tsujikawa 2013).
3 Alleviating the severe difficulty to detect if the fluid responsible for the cosmic speed up is under the form of a pure cosmological constant or is a time-dependent dark energy one.
the so-called circularity problem arises, i.e., the calibration issue between radiated energy or luminosity and the spectral properties that becomes plausible only if a background cosmology is a priori imposed\(^4\). Consequently, in view of the great uncertainty surrounding a model-independent procedure to calibrate GRBs, finding out new correlations that are background-independent becomes a crucial step to heal circularity.

In this paper, we assume the widely-consolidated model-independent technique of GRB calibration constructed by means of Bézier polynomials (Luongo & Muccino 2021b) and we apply it to the \(E_p - E_{iso}\) or Amati correlation (see e.g., Amati et al. 2008; Amati & Della Valle 2013), in particular to one of the best data set composed of 118 GRBs (Khadka et al. 2021). We propose a calibration procedure involving intermediate redshift catalogs, in lieu of \(z \approx 0\) data points, more often developed in the literature. We employ cosmic chronometers or OHD (see Capozziello et al. 2018, and references therein), and the most recent measurements of BAO (see Cao et al. 2021, and references therein). To do so, we fit both OHD and BAO data in conjunction by means of two Bézier parametric curves. These approximated curves are interconnected since BAO measurements contain information on \(H(z)\), without assuming an a priori hypothesis on the universe spatial curvature. Through this calibration procedure, the above GRBs can be viewed as standardized objects and can be used to test the standard spatially flat ΛCDM model and its minimal extension adding a non-zero spatial curvature parameter \(\Omega_k\). We thus investigate the effects of \(\Omega_k \neq 0\) and fix constraints over the free parameters by means of Markov chain – Monte Carlo (MCMC) analyses. The obtained results are not perfectly compatible with the current expectations since they roughly differ from Planck results on \(\Omega_k\). We investigate the corresponding systematics and show tighter bounds over the mass density, quite more similar to those found using standard candles. Further, we demonstrate that both spatial curvature and the mass density, quite more similar to those found using standard candles.

2 THEORETICAL WARM-UP

The mostly-adopted and investigated GRB correlation in the literature is built up through the rest-frame peak energy \(E_p\) of the \(\gamma\)-ray time-integrated \(\nu F_\nu\) energy spectrum and the isotropic energy \(E_{iso}\) radiated in \(\gamma\)-rays, i.e.,

\[
E_{iso} \equiv 4\pi D_L^2 S_0 (1+z)^{-1},
\]

(1)

where the observed bolometric GRB fluence \(S_0\) is evaluated from the integral of the \(\nu F_\nu\) spectrum in the rest-frame \(1-10^4\) keV energy band. Finally, a correction factor, namely \((1+z)^{-1}\), is involved in order to take into account cosmological redshift effects, i.e., to transform the inferred GRB overall duration into the source cosmological rest-frame measurements.

In this respect, it appears obvious that using GRBs with the aim of fitting cosmic data may be strongly affected by a few uncertainties, caused by selection and instrumental effects.

2.1 Building up the correlation

In view of the above considerations, the Amati correlation (Amati et al. 2002, 2008; Amati & Della Valle 2013; Demianski et al. 2017a; Dainotti & Amati 2018), typically dubbed \(E_p - E_{iso}\) relation, easily writes

\[
\log \left( \frac{E_p}{\text{keV}} \right) = a_0 + a_1 \log \left( \frac{E_{iso}}{10^{52}\text{erg}} \right),
\]

(2)

where the functional dependence requires an intercept \(a_0\) and a slope \(a_1\). In addition, we need to fix the dispersion \(\sigma_{iso}\) (D’Agostini 2005) and all the latter free terms require calibration.

We immediately see the caveat in Eq. (1), there \(E_{iso}\) depends on the background, i.e., the Hubble rate might be known a priori and so the corresponding dependence on the luminosity distance \(D_L\) is unavoidable.

2.2 GRB data set

To fulfill our fits, adopting the Amati relation, we here employ the most recent and largest data set of 118 bursts fulfilling the Amati correlation itself. For these data points, we underline they provide the smallest intrinsic dispersion (Khadka et al. 2021).

Moreover, to get model-independent cosmological bounds, in the following we now need to calibrate the relation by means of model-independent techniques, as below reported. Before that, we introduce the intermediate redshift data sets through which we intend to calibrate our GRB data set.

2.3 Intermediate redshift data sets

To calibrate GRB data sets we employ

– OHD, consisting at present of 31 measurements of the Hubble rate \(H(z)\). The corresponding measures are got at different redshifts (see, e.g., Capozziello et al. 2018).

– BAO, consisting of 15 measurements, split into 9 uncorrelated and 6 correlated, see Table 1.

For the sake of clearness, we focus on BAO data. The latter are often provided as volume-averaged distances \(D_{\text{clus}}^{\text{obs}} = \delta (r_s / r_{\text{fid}})\), where \(r_s\) is the comoving sound horizon at the baryon-drag epoch and \(r_{\text{fid}}\) is the value of \(r_s\) for the fiducial cosmological model used to convert redshift to distances. For each measurement, \(\delta\) and

\[
\gamma \equiv (r_s / r_{\text{fid}})^{\delta},
\]

(3)

are given in the third and fourth columns of Table 1, respectively.

The ratio \(\gamma\) has been computed by using the fiducial values \(r_{\text{fid}}\) provided in the references listed in Table 1. In particular, the value of \(r_s\) is taken from Planck Collaboration (2020).

Hence, by construction, BAO measurements provide

\[
D_V(z) = \left[ \frac{cz}{(1+z)^2} \frac{D_L^2(z)}{H(z)} \right]^{1/3},
\]

(4)

showing constraints on the cosmological parameters through their influence on the combined action of \(r_s\), \(H(z)\) and \(D_L(z)\).
Hence, as shown in Eq. (4), the only possibility to obtain constraints on \( D_L(z) \) without imposing a \textit{a priori} spatial curvature consists in using only BAO data points in conjunction with OHD. We will follow this more reliable strategy in what follows.

\[ H_n(x) = \sum_{i=0}^{n} g_{\alpha i} h_{\alpha}^{n}(x) , \quad h_{\alpha}^{n}(x) \equiv n! \frac{x^{i} (1-x)^{n-i}}{i! (n-i)!} . \quad (5) \]

where \( \alpha_i \) are the coefficients of the linear combination of the polynomials \( h_{\alpha}^{n}(x) \), re-scaled by a conventional factor \( g_n = 100 \text{ km/s/Mpc} \), being positive-defined for \( 0 \leq x \equiv z/z_{\text{max}}^{\text{OHD}} \leq 1 \), with \( z_{\text{max}}^{\text{OHD}} \), representing the maximum redshift of the OHD catalog. As proved in Luongo & Muccino (2021b), the only possible non-linear monotonic growing function over the redshift range of OHD has order \( n = 2 \), i.e., \( H_2(z) \). Moreover, by construction, it is possible to identify \( \alpha_0 = 0 \) with \( h_0 = H_0/(100 \text{ km/s/Mpc}) \).

BAO data \( \delta \) can be fitted by using Eq. (4) and the function \( H(z) \) can be approximated with \( H_2(z) \), extrapolated up to the BAO maximum redshift \( z_{\text{max}}^{\text{BAO}} \). In so doing, it is therefore licit to use BAO measurements in order to fit \( D_L(z) \) in a cosmology-independent way, again, by resorting a Bézier curve of order \( m \).

\[ D_m^2(y) = \sum_{j=0}^{m} g_{\beta j} d_m^j(y) , \quad d_m^j(y) \equiv m! \frac{y^{j} (1-y)^{m-j}}{j! (m-j)!} . \quad (6) \]

where \( \beta_j \) are the coefficients of the linear combination of the polynomials \( d_m^j(y) \) rescaled by a factor \( g_m = 1 \text{ Gpc}^2 \) and positive-defined for \( 0 \leq y \equiv z/z_{\text{max}}^{\text{OHD}} \leq 1 \). Immediately, from the above construction one argues how to extend the use of Bézier curves to BAO data points and, so, one highlights that:

\- by definition of cosmic distance, we need that \( D_m^2(y) \equiv 0 \) and so we are forced to start our Bézier expansion with \( j = 1 \);
\- the only non-linear monotonic growing function with the redshift has order \( m = 3 \), i.e., \( D_3^2(z) \) since \( j \) runs from 1 to 3.

Last but not least, it is remarkable to stress that the above determination of the luminosity distance is quite general. In other words, it includes information on the Universe’s spatial curvature, without introducing any theoretical bias, jeopardizing the overall picture and leading to circularity.

We are now ready to calibrate our free coefficients and to introduce the statical methods of numerical analyses, as we prompt below.

\subsection{3.2 Calibrating the coefficients with nested likelihoods}

We now estimate the coefficients \( \alpha_i \), \( 0 \leq i \leq 2 \), and \( \beta_j \), \( 1 \leq j \leq 3 \) through a nested likelihood approach. This method combines

\[ \text{model-independent} \]

But, in mind all the above ingredients, our prescription resides in interpolating OHD and BAO data sets without acquiring any \textit{cosmological model}, namely, to involve a model-independent calibration technique.\footnote{The words \textit{model-independent} typically rely on expansions and/or re-constructions of cosmic quantities, \textit{i.e.}, without any need of postulating a cosmological model, see e.g. (Aviles et al. 2012; Dunsby & Luongo 2016).} The here-employed method is attained by working out the so-called Bézier \textit{parametric curves}, obtained as a linear combination of Bernstein basis polynomials and firstly proposed in GRB contexts in Amati et al. (2019). The main steps to follow are thus summarized below.

\subsection{3.1 Bézier polynomials and GRB calibration}

The most general Bézier curve with established order \( n \), constructed by means of OHD data, can be written by

\[ D_m^2(y) = \sum_{j=0}^{m} g_{\beta j} d_m^j(y) , \quad d_m^j(y) \equiv m! \frac{y^{j} (1-y)^{m-j}}{j! (m-j)!} . \quad (6) \]
1) a fit of the OHD data in the range $0 \leq z \leq z_{\text{max}}^\text{OHD}$; and

2) a fit of the BAO data in the range $0 \leq z \leq z_{\text{max}}^\text{BAO}$, using also the extrapolation of $H_2(z)$ up to $z_{\text{BAO}} = 2$ by defining the function

$$\Delta \nu(z) = \left[ \frac{c z}{(1 + z)^2} \frac{D_L(z)^2}{H_2(z)} \right]^{1/3}. \quad (7)$$

Assuming Gaussian distributed errors, the total log-likelihood function of the model-independent estimate of $D_L(z)$ is given by

$$\ln \mathcal{L}_D = \ln \mathcal{L}_O + \ln \mathcal{L}_B, \quad (8)$$

where each contribution is described in details below.

- For OHD the log-likelihood function reads as

$$\ln \mathcal{L}_O = - \frac{1}{2} \sum_{k=1}^{N_O} \left\{ \frac{H_k - H_2(z_k)}{\sigma H_k} \right\}^2 + \ln(2\pi \sigma H_k^2), \quad (9)$$

where $N_O$ is the size of the OHD catalog with values $H_k$ and attached errors $\sigma H_k$.

- For BAO the log-likelihood function is given by

$$\ln \mathcal{L}_B = - \frac{1}{2} \sum_{k=1}^{N_B} \left\{ \frac{\delta_k - \Delta \nu(z_k)}{\sigma \delta_k} \right\}^2 + \ln(2\pi \sigma \delta_k^2), \quad (10)$$

where $N_B$ is the size of the BAO catalog with values $\delta_k$ and attached errors $\sigma \delta_k$.

The best-fit curves approximating both OHD and BAO catalogs and the resulting trend of the luminosity distance are portrayed in Fig. 1, where a comparison with the predictions of the $\Lambda$CDM paradigm (Planck Collaboration 2020) are also shown. The best-fit coefficients $\alpha$ and $\beta$, on which the plots in Fig. 1 are based, are displayed in the contour plots of Fig. 2 and summarized in Table 2.

Next, to get numerical bounds on the cosmological parameters, we employ the reconstructed luminosity distance portrayed in Fig. 1 to calibrate the Amati correlation in a model-independent as

$$E_{\text{iso}}^{\text{cal}}(z) \equiv 4 \pi D_L(z) S_8 (1 + z)^{-1}, \quad (11)$$

$$\sigma E_{\text{iso}}^{\text{cal}}(z) \equiv E_{\text{iso}}^{\text{cal}}(z) \left[ \frac{2 \sigma D_L(z)}{D_L(z)} \right]^2 + \left( \frac{\sigma S_8}{S_8} \right)^2, \quad (12)$$

where the error $\sigma E_{\text{iso}}^{\text{cal}}$ depends upon the errors on the GRB observable $S_8$, and the reconstructed luminosity distance $D_L(z)$.

To successfully calibrate all GRBs in each catalog, one needs to extrapolate $D_L(z)$ at redshifts higher than those of OHD and BAO catalogs. This, in principle may add further bias in the estimate of the cosmological parameters. To check this possibility, we performed a nested likelihood approach (Luongo & Maccino 2021b) that combines two sub-models involving:

i) a calibrator sample composed of GRBs in the range $0 \leq z \leq z_{\text{max}}^\text{EAE}$ (encompassing both OHD and BAO observations), employed for estimating the GRB correlation parameters, and

ii) a cosmological sample, i.e., the whole GRB data set, used to estimate the free model parameters.

Again, assuming Gaussian distributed errors, the total GRB log-likelihood function is given by

$$\ln \mathcal{L}_G = \ln \mathcal{L}_G^{\text{cal}} + \ln \mathcal{L}_G^{\text{cos}}. \quad (13)$$

The calibration log-likelihood is given by

$$\ln \mathcal{L}_G^{\text{cal}} = - \frac{1}{2} \sum_{k=1}^{N_{\text{cal}}} \left\{ \frac{Y_k - Y(z_k)}{\sigma Y_k} \right\}^2 + \ln(2\pi \sigma Y_k^2), \quad (14)$$

where $N_{\text{cal}}$ = 65 and

$$Y_k \equiv \log E_{p,k}, \quad (15a)$$

$$Y(z_k) \equiv a_0 + a_1 \left[ E_{\text{iso}}^{\text{cal}}(z_k) - 52 \right], \quad (15b)$$

$$\sigma Y_k^2 \equiv \sigma \log E_{p,k}^2 + a_1^2 \left[ \sigma \log E_{\text{iso}}^{\text{cal}}(z_k) \right]^2 + \sigma_a^2. \quad (15c)$$

The cosmological log-likelihood is given by

$$\ln \mathcal{L}_G^{\text{cos}} = - \frac{1}{2} \sum_{k=1}^{N_{\text{cos}}} \left\{ \frac{\mu_k - \mu_{\text{th}}(z_k)}{\sigma \mu_k} \right\}^2 + \ln(2\pi \sigma \mu_k^2), \quad (16)$$

where we have $N_{\text{cos}}$ = 118 and

$$\mu_k \equiv \frac{5}{2a_1} \left[ \log E_{p,k} - a_0 - a_1 \log \left( \frac{4\pi S_{th,k}}{1 + z_k} \right) \right], \quad (17a)$$

$$\sigma \mu_k^2 \equiv \frac{25}{4a_1^2} \left[ \sigma \log E_{p,k}^2 + a_1^2 \sigma \log S_{th,k}^2 + \sigma_a^2 \right]. \quad (17b)$$

3.3 Cosmic background and bounds over Bézier fits

We are now in position to experimentally test the non-flat $\Lambda$CDM paradigm by fixing it within $\mu_{\text{th}}(z)$ from Eq. (16). The corresponding Hubble rate becomes

$$H(z) = H_0 \sqrt{\Omega_m (1 + z)^3 + \Omega_\Lambda (1 + z)^2 + \Omega_k}, \quad (18)$$

where $\Omega_m$, $\Omega_k$, and $\Omega_\Lambda = 1 - \Omega_m - \Omega_k$ are matter, curvature, and cosmological constant density parameters, respectively. The luminosity distance (see, e.g., Goobar & Perlmutter 1993) is

$$D_L(z) = \frac{c}{H_0 \sqrt{\Omega_k}} S_k \sqrt{\frac{1}{H_k} \int_0^z H_0 d\tilde{z}^\prime}, \quad (19)$$

where $S_k(x) = \sinh(x)$ for $\Omega_k > 0$, $S_k(x) = x$ for $\Omega_k = 0$, and $S_k(x) = \sin(x)$ for $\Omega_k < 0$, and the distance modulus is

$$\mu_{\text{th}}(z) = 25 + 5 \log \left( \frac{D_L(z)}{\text{Mpc}} \right). \quad (20)$$

Finally, we perform MCMC fittings by searching for the best-fit parameters maximizing the log-likelihood defined in Eq. (13) and the 1 and 2-$\sigma$ contours. To do so, we modified the Wolfram Mathematica code from Arjona et al. (2019).

Depending on $r$, we decided to fix $h_0$ with the values of $\alpha_0$ obtained from the Bézier fitting, namely:

$$h_0 = 0.691^{+0.057}_{-0.032} \quad \text{for } r = 1, \quad (21a)$$

$$h_0 = 0.714^{+0.032}_{-0.044} \quad \text{for } r \neq 1. \quad (21b)$$

So, we finally explore both cases with free $\Omega_k$ and $\Omega_\Lambda \equiv 0$

The results are summarized in Table 3.

4 THEORETICAL INTERPRETATION OF NUMERICAL RESULTS

The above-developed strategy definitely shows that the GRB calibration can be performed adopting intermediate data points with respect to previous efforts involving low-redshift catalogs. In other words, we demonstrate that the need of more low-redshift GRBs is not fully-motivated in order to get feasible strategies of GRB
Figure 1. Plots of the best-fitting Bézier curves (blue thick lines) approximating $H(z)$, $D_V(z)$, and $D_L(z)$ with the 1–σ confidence bands (gray shaded areas): left for $r = 1$ and right for $r \neq 1$. A comparison with the ΛCDM paradigm (Planck Collaboration 2020) is also shown (see dashed red curves).

| OHD          | BAO          |
|--------------|--------------|
| $r = 1$      | $r \neq 1$   |
| $a_0$        | $a_1$        | $a_3$        | $\beta_1$        | $\beta_2$        | $\beta_3$        |
| $H(z)$       | $D_V(z)$     | $D_L(z)$     | $H(z)$           | $D_V(z)$          | $D_L(z)$          |
| $0.691_{-0.033}^{+0.032}$ | $0.979_{-0.123}^{+0.084}$ | $2.14_{-0.157}^{+0.181}$ | $-0.064_{-0.09}^{+0.099}$ | $42.071_{-8.145}^{+9.241}$ | $397.469_{-37.968}^{+44.243}$ |

Table 2. Best-fit Bézier coefficients $a_i$ of $H(z)$ and $\beta_j$ of $D_L^2(z)$ for $r = 1$ and $r \neq 1$.

| OHD          | BAO          |
|--------------|--------------|
| $r = 1$      | $r \neq 1$   |
| $a_0$        | $a_1$        | $a_3$        | $\sigma_{a_0}$ | $\Omega_0$ | $\Omega_b$ |
| $H(z)$       | $D_V(z)$     | $D_L(z)$     |               |             |             |
| $1.82_{-0.05}^{+0.10}$ | $0.70_{-0.03}^{+0.07}$ | $0.28_{-0.02}^{+0.09}$ | $0.25_{-0.01}^{+0.07}$ | $0.25_{-0.01}^{+0.07}$ | $0.25_{-0.01}^{+0.07}$ |
| $1.82_{-0.05}^{+0.10}$ | $0.70_{-0.03}^{+0.07}$ | $0.28_{-0.02}^{+0.09}$ | $0.25_{-0.01}^{+0.07}$ | $0.25_{-0.01}^{+0.07}$ | $0.25_{-0.01}^{+0.07}$ |

Table 3. Nested log-likelihood best-fit results and 1–σ (2–σ) errors for flat and non-flat ΛCDM models and for $r = 1$ and $r \neq 1$.

calibration that would fix the circularity issue. In this respect, we displayed in Figs. 1 suitable matching between theoretical and reconstructed curves. The case $r \neq 1$ is slightly less predictive than the one with $r = 1$, i.e., indicating that no particular deviations may occur as outputs of our MCMC analyses. This may be seen as a direct consequence of data reported in Table 1, where $r$ is effectively close to unity. For both the cases, we highlight two normalized Hubble constants got from Eqs. (21a). It is intriguing to notice that the case $r \neq 1$ provides a weakly larger value of $h_0$, albeit both cannot solve, in view of 2σ confidence level the Hubble tension today. Motivated by previous results toward this direction (Khadka et al. 2021; Khadka & Ratra 2020), it seems evident that,
even including GRBs into the computations of cosmic quantities, the tension remains unsolved. This conundrum could be therefore healed by switching to minimal extensions of the standard cosmological model (Izzo et al. 2012; Muccino et al. 2021; Luongo et al. 2022) or by additional terms within the Hilbert-Einstein action.

Moving to our experimental findings, we notice from Figs. 3 and 4 the remarkable fact the mass density is well constrained at the 1σ confidence level with respect to previous efforts making use of GRBs in the literature (Cao et al. 2022a,b). This certifies the goodness of intermediate calibrating techniques, demonstrating that our treatment is promising in refining cosmic outcomes by adopting GRBs. However, at 2σ confidence level the matter density appears weakly constrained, indicating that systematics may afflict the corresponding measurements. In general, r = 1 measurements appear surprisingly better bounded than free r parameters. In this case the values of Ω_m turn out to be closer to the Planck measurements in

Figure 2. Contour plots of the best-fit Bézier coefficients α_i of H(z) and β_j of D_L(z): left for r = 1 and right for r ≠ 1. Darker (lighter) regions mark the 1–σ (2–σ) confidence regions.

Figure 3. Contour plots of the best-fit correlation and cosmological parameters (for a ΛCDM model with Ω_k ≠ 0): left for r = 1 and right for r ≠ 1. Darker (lighter) regions mark the 1–σ (2–σ) confidence regions.
the case of spatially flat cosmology. For the sake of completeness, the same happens even for $r \neq 1$, despite the non-flat case provides values of $\Omega_m$ completely outside any theoretical expectations. It is significant to stress that spatial curvature cannot be constrained properly even at $1\sigma$ confidence level, leaving open the task to fix it at intermediate and high redshifts. In particular, the values of spatial curvature is only slightly compatible with zero, being just approximately on the left error bar for both our measurements. All the remaining free parameters appear to be well constrained at both the $1\sigma$ and $2\sigma$ confidence levels.

In view of our findings, we conclude the background cosmological model, namely the $\Lambda$CDM paradigm, is much more constrained by calibrating GRBs at intermediate redshifts. No significant expected departures in terms of dark energy are thus expected, but rather a plausible refinement of our fitting procedure, adding more data, could be useful to increase the values of matter densities here measured. A non-flat $\Lambda$CDM model is however debated and likely less probable than the flat case.

5 FINAL OUTLOOKS AND PERSPECTIVES

In this paper, we proposed how to calibrate GRBs in a model-independent way, adopting since the very beginning intermediate-redshift data sets, instead of low-redshift catalogs as commonly performed in the literature. To do so, we employed the novel calibration technique that makes use of Bézier polynomials, applying its use to the $E_{iso}-E_{kin}$ or Amati correlation. In this respect, we adopted one of the best data set composed of 118 GRBs and we worked out OHD and BAO surveys of data in conjunction by means of two Bézier parametric curves that allowed us to leave a priori free the spatial curvature. Hence, standardizing GRBs through the above technique, we tested the standard spatially flat $\Lambda$CDM model and its minimal extension adding a non-zero spatial curvature parameter, $\Omega_k$.

Hence, MCMC analyses have been performed by means of the Metropolis algorithm in two main cases: $r = 1$ and $r \neq 1$, showing findings that were not perfectly compatible with the current expectations since they roughly differ from Planck results on $\Omega_k$. Thus, in our four cases, namely $r = 1$ and $r \neq 1$ with either zero or non-zero curvature, we provided tight bounds over the mass density, quite more similar to those found using standard candles. Further, we investigated the corresponding systematics plaguing the approach and we underlined that, although our outcomes looked more suitable than previous results got from the literature, the $H_0$ tension was not addressed even with the use of GRBs. In analogy, the values of $\Omega_k$ were not bounded enough, but only slightly close to the zero value got from Planck measurements. Future works will focus on the use of refined Bézier model-independent curves and their use adopting further intermediate-redshift data points. Since including intermediate data points would refined $\Omega_m$, we will see how the calibration can be made by both low and intermediate data at the same time.

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