The Public Health Agency of Canada reported that nearly a quarter of all Canadians are obese (body mass index [BMI] > 30 kg/m²) at a direct cost of up to $4.6 billion per year. Obesity is associated with a reduction in quality of life and decreased life expectancy, and is linked to an increased incidence of diseases such as coronary heart disease, diabetes and certain types of cancers. However, bariatric surgery is recognized as a cost-effective treatment for severe obesity, resulting in sustained weight loss associated with the prevention, alleviation and resolution of many comorbid conditions.

Given the effectiveness of bariatric surgery for severe obesity and rising demand for surgical intervention, in 2009, the Ontario Ministry of Health and Long-Term Care invested $75 million to increase the number of surgeries performed in the province from 244 to 1470 per year by 2012. Four Bariatric Centres of Excellence were established that deliver care based on a multidisciplinary, multistage presurgical assessment process to appropriately identify suitable candidates for surgery. However, this model of care has led to prolonged preoperative evaluation, resulting in wait times that far exceed the provincially mandated target of 365 days from referral to surgery. The consequences of delayed surgery may include increased patient attrition, decreased patient satisfaction and delayed improvement in obesity-related comorbidities. Although a comprehensive evaluation is necessary and an extended preoperative course may be important to facilitate patient education and commitment, patients may also feel...

Patient and operational factors affecting wait times in a bariatric surgery program in Toronto: a retrospective cohort study

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Abstract

Background: Increasing rates of obesity have led to growing demand for bariatric surgery. This has implications for wait times, particularly in publicly funded programs. This study examined the impact of patient and operational factors on wait times in a multidisciplinary bariatric surgery program.

Methods: A retrospective study was conducted involving patients who were referred to a tertiary care centre (University Health Network, Toronto Western Hospital, Toronto) for bariatric surgery between June 2008 and July 2011. Patient characteristics, dates of clinical assessments and records describing operational changes were collected. Univariable analysis and multivariable log-linear and parametric time-to-event regressions were performed to determine whether patient and operational covariates were associated with the wait time for bariatric surgery (i.e., length of preoperative evaluation).

Results: Of the 1664 patients included in the analysis, 724 underwent surgery with a mean wait time of 440 (standard deviation 198) days and a median wait time of 445 (interquartile range 298–533) days. Wait times ranged from 3 months to 4 years. Univariable and multivariable analyses showed that patients with active substance use (β = 0.3482, p = 0.02) and individuals who entered the program in more recent operational periods (β = 0.2028, p < 0.001) had longer wait times. Additionally, the median time-to-surgery increased over 3 discrete operational periods (characterized by specific program changes related to scheduling and staffing levels, and varying referral rates and defined surgical targets; p < 0.001).

Interpretation: Some patients could be identified at referral as being at risk for longer wait times. We also found that previous operational decisions significantly increased the wait time in the program since its inception. Therefore, careful consideration must be devoted to process-level decision-making for multistage bariatric surgical programs, because managerial and procedural changes can affect timely access to treatment.

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discomfort with a longer wait time for surgery. In contrast to previous survey-based research on wait times for bariatric surgery,\textsuperscript{18-21} this study quantified how patient characteristics and operational factors contribute to excessive wait times in a large bariatric surgery program.

Methods

**Study design, setting and participants**

We conducted a retrospective review involving patients referred for bariatric surgery at a tertiary care centre (University Health Network, Toronto Western Hospital, Toronto), between June 1, 2008, and July 31, 2011. Only patients who underwent bariatric surgery or withdrew from the program at the time of data collection were included. We identified a total of 1682 eligible participants met the inclusion criteria. As a multidisciplinary, multistage program, the preoperative evaluation includes medical, social work, dietary, psychological or psychiatric assessments, and a surgical consultation before patients are deemed eligible for surgery (Figure 1).\textsuperscript{15} If substantial clinical or psychosocial issues are identified, an interdisciplinary team reviews each patient’s case to determine if intervention is appropriate or if they should be refused surgery. Patients must be cleared at each stage before moving on in the program. All physician referrals are administered through the Ontario Bariatric Network and distributed among regionalized surgical programs in the province. Selection criteria for referral are standardized across all Ontario bariatric surgical centres.\textsuperscript{22} All referred patients had a BMI greater than 40 kg/m\textsuperscript{2} or a BMI greater than 35 kg/m\textsuperscript{2} with at least 1 obesity-related comorbidity.

**Operational periods**

Several procedural changes took place during the study period (Table 1). We distinguished between 3 operational phases, because many of these interventions occurred at once. Each operational period was characterized by a unique set of managerial and administrative changes, related to fluctuations in staffing levels, referral rate, surgical target, and alterations in scheduling practices and new internal policy. The first period corresponds to early program performance (baseline levels), and the third period represents the most recent operational landscape. Staffing levels and the physical space reserved for assessments increased in each period since the program’s inception.

**Data collection and outcome measures**

Patient characteristics (height, weight and BMI) including demographic information (age, sex and postal code), possible exclusion criteria for surgery (substance use including smoking, alcohol and drug use) and the dates of each assessment including surgery were collected from the referral document and electronic patient records. We used a broad definition for substance use because any substance use is a contraindication for surgery and must be stopped before a patient can have surgery. Data were linked to operational records describing procedural changes. The referral dates for 5 patients and the physical charts for 13 patients were missing, and, therefore, they were excluded from the analysis.

The primary outcome was the overall wait time for bariatric surgery. This represents patients who completed all preoperative assessments. We also examined the total wait time for patients who attended an orientation session, therefore focusing on patients who showed a willingness to participate in the program. Patients who attended orientation but did not undergo surgery were censored as of their last appointment date.

**Statistical analysis**

We used the R programming environment for data analysis. A random forest technique was used to impute missing data.\textsuperscript{23} Group comparisons were made using univariable, nonparametric statistical testing. As recommended in the STROBE statement,\textsuperscript{24} BMI, age and distance were modelled as both continuous and categorical variables. We used multivariable regression analysis to determine the association between patient and operational covariates and total wait time for all patients who underwent bariatric surgery (log-linear) and all patients who attended orientation but who did not necessarily undergo surgery (parametric time-to-event). Departures from linearity were assessed by plotting a locally weighted scatter plot curve through the Martingale residuals. We tested the appropriateness of the regression models by performing a global validation of model assumptions, a variance inflation test for multicollinearity, a Breusch–Pagan test for heteroskedasticity and a

![Figure 1: Patient flow through the bariatric surgery program.](image-url)
Q-Q plot of the residuals for normality. A 95% confidence interval (CI) was used to assess statistical significance.

The study protocol was approved by the University Health Network Research Ethics Board.

Results

Univariable analysis

Of the 1664 patients included in the analysis, 724 underwent surgery with a mean wait time of 440 SD 198) days and a median wait time of 445 (interquartile range [IQR] 298–533) days. Wait times ranged from 3 months to 4 years. The majority of referred patients (74.4%) and those who had surgery (80.8%) were female. Patients who did not reach surgery either dropped out of the program (self-removal) or were refused surgery because of substantial clinical or psychosocial issues (removal from the program on the advice of the interdisciplinary team). The average age at referral was 48 (SD 11.3) years, and few patients (n = 42) with active substance use were referred to the program. The median patient distance to the hospital was 48.2 km (range 1–550 km). The values for the referral BMI of 83 patients were missing; therefore, we employed a nonparametric multiple imputation procedure to provide estimates of BMI for those patients. The out-of-bag error estimate was 3.6%.

Differences in wait time for surgery by patient characteristic and operational period are shown in Table 2. We found no difference in the length of the preoperative assessment among patients by age (χ² = 6.538, df = 4, p = 0.2), BMI (χ² = 3.186, df = 3, p = 0.4) and by distance from the hospital (χ² = 6.256, df = 6, p = 0.4). However, subsequent analysis using the Wilcoxon rank sum test (also known as the Mann–Whitney test) showed that the median distance to hospital for patients who attended multiple assessments in 1 day was 30 km farther than the distance for patients who did not have this type of assessment schedule (W = 36 780.6, p < 0.001). Males (W = 36 433.5, p = 0.042), individuals with active substance use (W = 1596.5, p = 0.009) and patients who attended an orientation session in Period 2 (W = 44 809, p < 0.001), experienced a longer overall wait time. Patients who had surgery in Period 3 spent more time in the program than in any other period (χ² = 160.8, df = 2, p < 0.001). Differences between the surgery and no-surgery cohorts are summarized in Table 3. A more detailed analysis of factors affecting attrition was presented in a previous study.17

Table 1: Operational changes that took place over the study period (June 1, 2008, to July 31, 2011)*

| Operational change | Period 1 | Period 2 | Period 3 |
|--------------------|---------|---------|---------|
| Staffing levels    | Baseline| †       | †       |
| Surgical target†   | 180     | 300     | 270     |
| Physician referral rate† | Baseline | †       | ⇔       |
| Internal procedures | —      | Follow-up scheduling‖ | No-show policy** |

*Adapted from Table 3 in Diamant et al. (2014).† Represents the maximum allowable number of surgeries that can be performed annually as defined by the Ontario provincial government.§Represents the average number of weekly referrals by primary care physicians.¶No change.‖Scheduling began to include follow-up appointments for patients who had undergone surgery.**Formal removal process initiated for no-show patients (i.e., patients who failed to attend 3 scheduled assessments).

Table 2: Differences in time-to-surgery* (only patients who underwent bariatric surgery included in the analysis (n = 724)*

| Characteristic | Median (IQR) | p value |
|---------------|-------------|---------|
| Sex           |             |         |
| Male          | 446 (327–580) | 0.04†   |
| Female        | 409 (293–525) |         |
| Substance use |             |         |
| Yes           | 557 (522–630) | 0.009†  |
| No            | 413 (294–532) |         |
| BMI, kg/m²    |             | 0.4     |
| 35–39         | 418 (275–605) |         |
| 40–49         | 412 (297–520) |         |
| 50–59         | 414 (293–540) |         |
| ≥ 60          | 444 (334–532) |         |
| Age, yr       |             | 0.2     |
| 19–29         | 414 (337–599) |         |
| 30–39         | 404 (307–527) |         |
| 40–49         | 410 (276–527) |         |
| 50–59         | 415 (293–515) |         |
| ≥ 60          | 473 (351–567) |         |
| Distance, km  |             | 0.4     |
| 0–25          | 405 (272–520) |         |
| 25–50         | 417 (271–533) |         |
| 50–100        | 422 (312–535) |         |
| 100–200       | 438 (307–558) |         |
| 200–300       | 398 (308–483) |         |
| 300–400       | 450 (370–525) |         |
| 400–500       | 418 (341–523) |         |
| > 500         | 474 (322–627) |         |
| Operational period (orientation) | < 0.001† |
| 1             | 369 (226–519) |         |
| 2             | 432 (335–539) |         |
| Operational period (surgery) | < 0.001† |
| 1             | 373 (164–514) |         |
| 2             | 381 (273–480) |         |
| 3             | 608 (482–784) |         |

Note: BMI = body mass index, IQR = interquartile range.
*The Mann–Whitney test (Kruskal–Wallis test) was used to determine whether 2 (3 or more) populations spent a similar amount of time in the program.
†p < 0.05 was considered statistically significant.
Multivariable analysis

The results of a log-linear regression of wait time on covariates known at referral are presented in Table 4. Only patients who underwent bariatric surgery were included in this analysis. Patients with active substance use ($\beta = 0.3482$, $p = 0.02$) and those who attended orientation in Period 2 ($\beta = 0.2028$, $p < 0.001$) spent more time in the program. A multivariable time-to-event regression for patients who attended an orientation session is shown in Table 5. Again, patients with active substance use ($\beta = 1.489$, $p = 0.024$) and those who attended orientation in Period 2 ($\beta = 1.279$, $p < 0.001$) spent more time in the program.

We conducted additional multivariable analyses using power-transformed regression to explore how covariates affected the wait time between 2 contiguous preoperative assessments (interstation wait time). Several regression models were estimated and they confirmed the results presented above (not shown). Specifically, active substance use and operational period were statistically significant in all models. Wait times increased with operational period and showed a worsening trend. Additionally, the time between referral and orientation was slightly reduced for males and older patients, whereas interstation wait time increased with the logarithm of distance from the hospital in the early stages of the program only.

The mean wait time between specific assessment stations (and by operational period) is presented in Table 6. The highlighted cells represent the expected sequence of appointments as determined by program directors. Although this is the modal pathway, many patients deviated from this sequence of care.

Interpretation

Our results showed that specific patient profiles and operational period are associated with longer wait times. We found that the time patients spent in the program did not depend on BMI and was generally insensitive to age, sex and distance from the bariatric centre. However, active substance use was associated with longer wait times, and patients who attended orientation in Period 2 or who had surgery in Period 3 had longer wait times. We also showed that process-level changes are associated with worsening wait times (i.e., since the program was established, the wait time has steadily increased despite several operational interventions). Furthermore, patients spent 75 days longer, on average, in the program than the government-mandated target for bariatric surgery. Whereas 46% of that time was spent between referral and orientation, 71% of all patient transitions between assessments took longer than 30 days. This is a problematic trend that could impact timely patient access to treatment.

Substance use was associated with longer preoperative evaluation and was shown to independently predict overall wait time. This may be partially explained by the 3-month abstinence requirement before patients can attend further preoperative assessments or the 6-month abstinence requirement before patients can proceed to surgery. Despite active substance use being an exclusion criterion for referral, more stringent practices may need to be adopted that would allow programs to refuse admittance to patients with uncontrolled substance dependencies. Our results showed that active sub-

| Characteristic                              | Surgery | No surgery | p value  |
|---------------------------------------------|---------|------------|----------|
| Sex                                         |         |            | < 0.001§ |
| Male                                        | 139     | 287        |          |
| Female                                      | 585     | 653        |          |
| Substance use                               |         |            | < 0.001§ |
| Yes                                         | 9       | 33         |          |
| No                                          | 715     | 907        |          |
| BMI, kg/m²; mean ± SD                       | 48.87 ± 8.19 | 47.67 ± 7.92 | 0.003§  |
| Age, yr; mean ± SD                          | 46.78 ± 10.54 | 48.42 ± 11.86 | 0.003§  |
| Distance from the hospital, km; mean ± SD   | 119.73 ± 179.49 | 113.64 ± 177.57 | 0.5     |
| Operational period (referral)†             |         |            | 0.047§   |
| 1                                           | 478     | 575        |          |
| 2                                           | 246     | 347        |          |
| Operational period (orientation)‡           |         |            | < 0.001§ |
| 1                                           | 233     | 155        |          |
| 2                                           | 490     | 507        |          |

Note: BMI = body mass index.

*The Pearson $\chi^2$ test and Student $t$ test were used as appropriate.
†Binary variable indicating if a patient was referred in Period 1 or Period 2.
‡Binary variable indicating if a patient attended orientation in Period 1 or Period 2.
§p < 0.05 was considered statistically significant.
stance users are still referred for surgery. Enforcing the policy to restrict program entry until patients can demonstrate pro-
longed abstinence or until they are referred again would help to
relieve system congestion and reduce wait times for pa-
patients identified as ideal surgical candidates. Moreover,
patients presenting with substance use should be identified at
the time of referral, because they are less likely to undergo
bariatric surgery. This would facilitate earlier behavioural
modification (e.g., smoking cessation) and delay immersion
into an already resource-constrained system. However,
restrictions at the provincial level may limit the autonomy
that programs and individual providers have in deciding
which patients should not be considered for surgery based on
efficacy and safety concerns. Conversely, living farther away
from the bariatric centre did not lead to increased wait times
for patients. Although distance may affect the wait time
between appointments, patients who lived farther away from
the hospital tended to cluster their appointments (i.e., sched-
ule multiple consecutive assessments on a single day) and
made fewer trips. This finding is supported by a previous
study that showed that distance did not appear to be a barrier
to care in this setting.Operational period, as characterized by specific managerial and
procedural changes, influenced wait times. Patients who
had surgery in Period 3 spent more time in the program than
in any other period. This suggests that process-level decision-
making that guided operational practices did not mitigate
increases in demand. During Period 3, the maximum available
weekly appointments were scheduled, and patients who had
already undergone bariatric surgery were being scheduled for
postsurgical follow-ups at several assessment stations. As the
number of surgeries increased, so did the number of patients
needing follow-up appointments, and, therefore, a smaller
pool of appointments were available to the increasing num-
bers of presurgical patients. This operational roadblock may
explain why the surgical target was subsequently lowered in
Period 3 from that in Period 2 despite an increase in the
referral rate. Finally, changes in internal operating proce-
dures, specifically the initiation of a formal removal process
for no-show patients, also may have contributed to reduced
operational performance. Although wait times could be
improved through initiatives to expand capacity, obtaining
additional resources is not possible within the current fund-
ing envelope. Novel scheduling techniques need to be intro-
duced to achieve optimization. For instance, preferential
scheduling practices or combined assessments could be used
to prioritize patients at low-risk for long wait times and non-
completion. These patients would be fast-tracked through
the program while more resources are directed toward
patients requiring intensive evaluation. Other interventions
can also be implemented to improve performance, such as
group counselling and enhanced triaging of patients. For
instance, if relevant medical and psychosocial information is
collected at referral, patients can be directed to assessments
that best address any potential clinical issues early in their
sequence of care.
This study represents the experience of a Canadian institu-
tion and a bariatric program operating within a public funding
model. It was limited to retrospective data and did not take
into account information that was learned by providers at pre-
surgical assessments. It has implications for other types of
multidisciplinary programs across different clinical specialties
that involve longitudinal assessment of patients in a complex

| Table 4: Differences in wait time for patients who underwent bariatric surgery (n = 724) |
|-------------------------------|-----------------|--------|--------|
| Covariates*                  | Estimate ± SE, d | t     | p value |
| (Intercept)                  | 334.7 ± 30.3    | 64.32  | 0.0000 |
| Substance use                | 139.4 ± 73.8    | 2.27   | 0.02†  |
| Age                         | −0.034 ± 0.54   | −0.08  | 0.9    |
| BMI, kg/m²                   |                 |       |        |
| 40–49                       | −11.93 ± 19.40  | −0.60  | 0.5    |
| 50–59                       | −2.10 ± 21.7    | −0.10  | 0.9    |
| 60+                         | −1.64 ± 26.48   | −0.06  | 0.9    |
| Male                        | 22.23 ± 15.69   | 1.47   | 0.1    |
| Log, distance               | 3.668 ± 4.142   | 0.89   | 0.4    |
| Operational period 2†       | 75.25 ± 16.73   | 5.44   | < 0.001‡ |

Note: BMI = body mass index.
*Baseline covariate values: female, no substance use, 19 years of age, BMI between 35 and 39 kg/m², lived within a few kilometres of the bariatric centre, and referred or attended orientation in Period 1.
†Binary variable indicating if a patient attended orientation in Period 1.
‡p < 0.05 was considered statistically significant.

| Table 5: Differences in wait time for patients who attended orientation (n = 1385)* |
|-------------------------------|-----------------|--------|--------|
| Distribution parameters      | Estimate ± SE, d |        |        |
| Mean                         | 431.80 ± 28.79  |        |        |
| SD                           | 1.6117 ± 0.195  |        |        |
| Covariates†                  | OR (95% CI)     | p value |
| Substance use                | 1.49 (1.144–1.937) | 0.02§  |
| Age                          | 1.00 (0.998–1.004) | 0.8    |
| BMI, kg/m²                   |                 |        |        |
| 40–49                        | 0.93 (0.834–1.030) | 0.05   |
| 50–59                        | 0.95 (0.849–1.070) | 0.3    |
| 60+                          | 0.93 (0.808–1.073) | 0.06   |
| Male                         | 1.09 (1.005–1.175) | 0.08   |
| Log, distance                | 0.98 (0.962–1.050) | 0.07   |
| Operational Period 2†        | 1.28 (1.194–1.370) | < 0.001§ |

Note: BMI = body mass index, CI = confidence interval, OR = odds ratio.
*Parametric time-to-event regression with log-normal hazards. Dependent variable was time-to-surgery. Patients who did not undergo bariatric surgery were censored at their last assessment date. Patients who did not attend orientation were excluded from the analysis.
†Baseline covariate values: female, no substance use, 19 years of age, BMI between 35 and 39 kg/m², lived within a few kilometres of the bariatric centre, and referred or attended orientation in Period 1.
‡Binary variable indicating if a patient attended orientation in Period 2.
§p < 0.05 was considered statistically significant.
system (e.g., transplant and oncology). Direct application of these findings may be limited, because bariatric surgery is an elective procedure that, if not performed urgently, does not have immediate patient consequences.

Our analysis has revealed 3 areas for improvement. First, certain types of patients should be identified early in the process, and program administrators may want to develop customized care plans that address their specific needs. Second, the current exclusion criteria for surgery may be too lenient, especially the criteria that relate to substance use. Third, previous operational interventions have not decreased wait times. We are currently investigating how to schedule patients for assessments to improve health care provider utilization, and what the optimal proportion of new to follow-up appointment slots should be. Furthermore, better triaging enables early treatment and relevant diagnostic tests (e.g., blood work, sleep study) to be completed even while patients undergo further assessments. This would decrease the waiting time when patients are closer to surgery and could also reduce late-stage patient attrition. Future research will identify triage techniques that improve wait times within this unique clinical setting. Additional inquiry into the effectiveness of a longitudinal care model and utility of a 1-year target wait time for measuring efficiency is also needed.

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