MOA-2011-BLG-322Lb: a “second generation survey” microlensing planet

Y. Shvartzvald1,* , D. Maoz1, S. Kaspi1, T. Sumi2,16, A. Udalski3,17, A. Gould4, D.P. Bennett5,16, C. Han18 and

F. Abe6, I.A. Bond7, C.S. Botzler8, M. Freeman8, A. Fukui9, D. Fukunaga6, Y. Itow6, N. Koshimoto2, C.H. Ling7, K. Masuda6, Y. Matsubara6, Y. Muraki6, S. Namba2, K. Ohnishi10, N.J. Rattenbury8, To. Saito11, D.J. Sullivan12, W.L. Sweatman7, D. Suzuki2, P.J. Tristram13, K. Wada2, P.C.M. Yock8,

(MOA collaboration),

and

J. Skowron3, S. Kozłowski3, M.K. Szymański3, M. Kubiaκ3, G. Pietrzyński3,14, I. Soszyński3, K. Ulaczyk3, L. Wyrzykowski3,15, R. Poleski3,4, P. Pietrukowicz3

(OGLE collaboration)

1 School of Physics and Astronomy, Tel-Aviv University, Tel-Aviv 69978, Israel
2 Department of Earth and Space Science, Osaka University, Osaka 560-0043, Japan
3 Warsaw University Observatory, Al. Ujazdowskie 4, 00-478 Warszawa, Poland
4 Department of Astronomy, Ohio State University, 140 West 18th Avenue, Columbus, OH 43210, United States of America
5 University of Notre Dame, Department of Physics, 225 Nieuwland Science Hall, Notre Dame, IN 46556-5701, USA
6 Solar-Terrestrial Environment Laboratory, Nagoya University, Nagoya, 464-8601, Japan
7 Institute of Information and Mathematical Sciences, Massey University, Private Bag 109-904, North Shore Mail Centre, Auckland, New Zealand
8 Department of Physics, University of Auckland, Private Bag 92-019, Auckland 1001, New Zealand
9 Okayama Astrophysical Observatory, National Astronomical Observatory of Japan, Asakuchi, Okayama 719-0232, Japan
10 Nagano National College of Technology, Nagano 381-8550, Japan
11 Tokyo Metropolitan College of Aeronautics, Tokyo 116-8523, Japan
12 School of Chemical and Physical Sciences, Victoria University, Wellington, New Zealand
13 Mt. John University Observatory, P.O. Box 56, Lake Tekapo 8770, New Zealand
14 Universidad de Concepción, Departamento de Astronomía, Casilla 160-C, Concepción, Chile
15 Institute of Astronomy, University of Cambridge, Madingley Road, Cambridge CB3 0HA, UK
16 Microlensing Observations in Astrophysics (MOA) Collaboration
17 Optical Gravitational Lens Experiment (OGLE) Collaboration
18 Department of Physics, Chungbuk National University, Cheongju 371-763, Republic of Korea

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ABSTRACT

Global ”second-generation” microlensing surveys aim to discover and characterize extrasolar planets and their frequency, by means of round-the-clock high-cadence monitoring of a large area of the Galactic bulge, in a controlled experiment. We report the discovery of a giant planet in microlensing event MOA-2011-BLG-322. This moderate-magnification event, which displays a clear anomaly induced by a second lensing mass, was inside the footprint of our second-generation microlensing survey, involving MOA, OGLE and the Wise Observatory. The event was observed by the survey groups, without prompting alerts that could have led to dedicated follow-up observations. Fitting a microlensing model to the data, we find that the timescale of the event was $t_E$ =23.2±0.8 days, and the mass ratio between the lens star and its companion is $q =0.028±0.001$. Finite-source effects are marginally detected, and upper limits on them help break some of the degeneracy in the system parameters. Using a Bayesian analysis that incorporates a Galactic structure model, we estimate the mass of the lens at $0.39^{+0.42}_{−0.19} M_J$, at a distance of $7.56±0.91$ kpc. Thus, the companion is likely a planet of mass $11.6^{+13.4}_{−5.6} M_J$, at a projected separation of $4.3^{+1.5}_{−1.2}$ AU, rather far beyond the snow line. This is the first pure-survey planet reported from a second-generation microlensing survey, and shows that survey data alone can be sufficient to characterize a planetary model. With the detection of additional survey-only planets, we will be able to constrain the frequency of extrasolar planets near their systems’ snow lines.

Key words: surveys – gravitational lensing: micro – binaries: general – planetary systems – Galaxy: stellar content
1 INTRODUCTION

The discovery of thousands of extrasolar planets ranks among the most exciting scientific developments of the past decade. The majority of those exoplanets were detected and characterized by the transit and radial velocity methods, which favor the detection of massive planets in close orbits around their hosts, stars at distances within a few hundred parsecs. Microlensing, in contrast, can reveal planets down to Earth mass and less (Bennett & Rhie 1994), at larger orbits - about 1 to 10 AU, which is where the “snowline” is located, and beyond which giant planets are expected to form according to planet formation models (Ida & Lin 2005). Microlensing enables the detection of planets around all types of stars at distances as far as the Galactic center, and even planets unbound from any host star (Sumi et al. 2011). Although planets discovered with microlensing still number in the few tens, several teams (e.g. Gould et al. 2010; Sumi et al. 2010; Cassan et al. 2012) have attempted to estimate the frequency of planets at these separations. Gould et al. (2010) estimated from high-magnification events a $\sim 1/6$ frequency of Solar-like systems. Sumi et al. (2011) found that Neptune-mass planets are three times more common than Jupiters beyond the snowline. Cassan et al. (2012), concluded that, on average, every star in the Galaxy hosts a snowline-region planet. Moreover, the fact that two out of the 20 planetary systems discovered by microlensing, OGLE-2006-BLG-109Lb,c (Gaudi et al. 2008) and OGLE-2012-BLG-0026Lb,c (Han et al. 2013), host 2 planets, suggests that multiple systems are common, as also indicated at smaller separations by transit data from Kepler (Howard 2013). Over the past decade, microlensing planet discoveries have largely come from observing campaigns in which specific, high-magnification (A $\gtrsim$ 100), events are followed intensively by networks of small telescopes, in order to detect and characterize planetary anomalies in the light curves. High magnification events are very sensitive to planets near their snowlines, but they are rare events. Furthermore, the inhomogeneous social process through which potentially high magnification events are alerted and followed up complicates their use for statistical inferences on planet frequency.

Microlensing surveys have been in transition into the so-called “second generation” phase (Gaudi et al. 2009), wherein a large area of the Galactic bulge is monitored continuously, with cadences high enough to detect planetary anomalies without any follow-up observations or changes to the observing sequence (e.g. switching to a higher cadence). The first such survey began in 2011 and combines three groups: OGLE (Optical Gravitational Lensing Experiment) – observing from Chile (Udalski 2003), MOA (Microlensing Observations in Astrophysics) – observing from New-Zealand (Sumi et al. 2003), and the Wise survey observing from Israel (Shvartzvald & Maoz 2012). Shvartzvald & Maoz (2012) have simulated the results that can be expected from this “controlled experiment”, and its potential to measure the abundance of planetary systems. They found that the overall planet detection efficiency for the survey is $\sim 20\%$. In the 2011 season there was a total of 80 events inside the high-cadence survey footprint that is common to all three groups (an additional 218 events were observed by only two of the groups). Of those 80 events, at least three showed a clear planetary anomaly: MOA-2011-BLG-293 (Yee et al. 2012), MOA-2011-BLG-322 (this paper) and OGLE-2011-BLG-0265 (in preparation). Accounting for our detection efficiency, these results imply $\sim 1/5$ frequency of planetary systems, which is in line with previous estimates by Gould et al. (2010).

In this paper, we present the analysis of MOA-2011-BLG-322Lb. This is the first planetary microlensing event that is detected and analyzed based solely on second-generation survey data. In principle, OGLE-2003-BLG-235/MAO-2003-BLG-53Lb (Bond et al. 2004) and MOA-2007-BLG-192Lb (Bennett et al. 2008) were also discovered and characterized based only on MOA and OGLE data. This was fortuitously possible despite the sparse sampling of first generation surveys. MOA-2011-BLG-293Lb (Yee et al. 2012) was also characterizable by survey-only data, but it did include a large amount of non-survey data that were prompted by alerts, following early realization of the event’s high magnification. We describe the observations by the three survey groups in Section 2. In Section 3 we present the binary microlensing model fitted to the event. A Bayesian analysis estimating the physical properties of the system is presented in Section 4 and we discuss our results in Section 5.

2 OBSERVATIONAL DATA

The microlensing event MOA-2011-BLG-322 was first detected on June 30, 2011, 16:15 UT by MOA, who operate the 1.8m MOA-II telescope at the Mt John Observatory in New Zealand. The source is located inside the second-generation survey footprint, at RA = 18:04:53.6, Dec = $-27\,^\circ\,13\,^\prime\,15.4\arcsec$ (J2000.0). Thus the event was also observed by the Wise team with the 1m telescope at the Wise Observatory in Israel, and by the OGLE team with the 1.3m Warsaw telescope at the Las Campanas Observatory in Chile. It was discovered independently by the OGLE early warning system (EWS, Udalski 2003) and designated as OGLE-2011-BLG-1127. Since the event was not identified as interesting in real time (although MOA noted some anomalous behavior), the survey teams continued their regular observing cadences for this field. The observational information for each group (filter, cadence, exposure time) is summarized in Table 1. The event was not observed by most of the microlensing follow-up groups such as RoboNet-III (Tsapras et al. 2004), MiNDSTEp (Dominik et al. 2010), or PLANET (Albrow et al. 1998) since it was not very bright at baseline, and had an apparently low magnification ($\sim 1$ mag, although we show in Section 3 below, that this is due to blending, and the true magnification was a moderate $A = 23$). We note that Farm Cove Observatory from $\mu$Fun (Gould 2008) tried to observe the event for a couple of nights after the anomaly, but the data were too noisy to be of use.

OGLE and MOA data were reduced by their standard difference image analysis (DIA) procedures (Bond et al. 2001; Udalski 2003). The Wise data were reduced using the pySis DIA software (Albrow et al. 2005). The MOA and Wise fluxes were aligned to the OGLE I-band magnitude scale, and inter-calibrated to the microlensing model (see Section 3). As further described below, this event includes a large amount of “blended light”, which could be
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Figure 1. Inter-calibrated light curve for the 3 datasets: MOA – green, OGLE – blue, Wise – red. The curve is our best-fit binary lens model (bottom panel shows the residuals from the model). The apparent magnification is only $\sim 1$ mag due to the large blending fraction. (b) Binned OGLE and MOA data showing a $\sim 0.02$ mag bump, 50 days after the peak of the event, which may be real, although similar features can be found over the light-curve baseline.

Table 1. Observational Summary

| Group | Filter | Cadence [min] | Exp. time [sec] |
|-------|--------|---------------|-----------------|
| MOA   | MOA-Red$^*$ | 18            | 60              |
| Wise  | $I$    | 30            | 180             |
| OGLE  | $I$    | 60            | 100             |

$^*$equivalent to $R + I$.

due to unrelated stars projected near the source and lens stars, and/or due to the lens star itself. Re-reduction of the pipeline-reduced data from each observatory, including centroid alignment and correction for a trend of DIA flux with seeing width for MOA and Wise data (seen for this source in baseline data, before the event), corrected a number of measurements with systematic errors in the observatories’ pipeline reductions.

Figure 1 shows the observed light curve, with a clear deviation from symmetric, point-mass microlensing (Paczynski 1986). We therefore proceed to the next level of complexity and attempt to model this event as a binary lens.

3 MICROLENSEING MODEL

The basic binary-lens microlensing model requires seven parameters. Three are the standard “Paczynski” parameters: $t_E$ – the event time scale, i.e. the time it takes the source to cross the Einstein radius of the primary mass; $u_0$ – the impact parameter, i.e. the minimum angular separation, in units of the Einstein angle, between the source and lens; and $t_0$ – the epoch of minimum separation. We note that, in our model, these three parameters are with respect to the primary lens star (and not, e.g., with respect to the center of mass of the lens and its companion). The Einstein radius is $\theta_E \cdot D_L$, where the Einstein angle is defined through

$$\theta_E^2 = \kappa M \pi_{\text{rel}}, \quad \kappa = \frac{4G}{c^2},$$

where $M$ is the mass of the lens star, and the relative parallax is

$$\pi_{\text{rel}} = \frac{1}{D_L} - \frac{1}{D_S},$$

where $D_S$ and $D_L$ are the distances to the source and lens stars, respectively.

The possibility of finite source effects due to the non-
Table 2. Microlensing model

| Parameter | Value |
|-----------|-------|
| $t_0$ [HJD] | 2455774.2729(54) |
| $u_0$ | 0.04686(17) |
| $t_E$ [days] | 23.17(80) |
| $s$ | 1.822(10) |
| $\alpha$ [rad] | 0.3662(36) |
| $q$ | 0.0284(10) |
| $\rho$ | < 0.007 |
| $f_s/f_b$ [OGLE] | 0.0573(20) |
| $I_{source}$ [mag] | 19.83(5) |
| $I_{blend}$ [mag] | 16.69(2) |
| $c_{moa}$ | 1.50 |
| $c_{wise}$ | 1.25 |
| $c_{ogle}$ | 1.57 |

The recovered probability distribution for the finite source size, setting an upper limit of $\rho < 0.007$, and showing that the data are consistent with a point source. The distribution is used to construct a prior for our Bayesian analysis.

of freedom (DOF) equals unity, using

$$\chi_0^2 = c_i \sqrt{\sigma_i^2 + \epsilon_{\text{min},i}}^2$$

$$\sigma_i' = c_i \sqrt{\sigma_i^2 + \epsilon_{\text{min},i}}'$$

where $\sigma_i$ is the pipeline-reported flux errors for each group, and $c$ and $\epsilon_{\text{min}}$ are the re-normalization coefficients. $\epsilon_{\text{min}}$ is a systematic photometric error floor that dominates when the source is very bright and the Poisson errors become small.

Figure 3. The recovered probability distribution for the finite source size, setting an upper limit of $\rho < 0.007$, and showing that the data are consistent with a point source. The distribution is used to construct a prior for our Bayesian analysis.

In addition to these seven physical parameters, there are two calibration parameters for each dataset $i$, representing the flux from the source star ($f_{s,i}$) and the blended flux $f_{b,i}$ from other stars in the line of sight, including the lens. For a true magnification $A(t)$ at time $t$, the observed flux $F(t)$ is therefore

$$F(t) = f_{s,i} \cdot A(t) + f_{b,i}$$

In order to solve for the magnification of the binary-lens model, we use the ray-shooting light-curve generator described in Shvartzvald & Mao (2012). Briefly, we construct a trial model of the binary lens with a given choice of the parameters in the problem. We divide the lens plane onto a grid and use the lens equation directly to map the lens plane onto the source plane. By calculating the entire magnification $A(t)$ at once, and by using an adaptive grid that increases the lens-plane resolution around the image positions, we achieve fast computation times. The magnification is then the ratio of the summed solid angles subtended by all the images in the lens plane to that of the solid angle subtended by the source.

A Markov-chain Monte-Carlo (MCMC) search of the parameter space, using Gibbs sampling, is used to find the best-fit solution and its uncertainty. For every trial model, the calibration parameters were found analytically by means of a linear least squares minimization. Following previous microlensing event analyses (see e.g. Yee et al. 2012), the flux errors, $\sigma_i'$, were re-normalized such that $\chi^2$ per degree of freedom (DOF) equals unity, using

$$\sigma_i' = c_i \sqrt{\sigma_i^2 + \epsilon_{\text{min},i}}'$$

The trajectory of the source relative to the caustic structure is shown in Figure 2. Since the source passes near the central caustic, we have checked for the so-called $s \leftrightarrow s^{-1}$ degeneracy, that often occurs with anomalies dominated by the central caustic in the case of high-magnification events (Griest & Safizadeh 1998). However, the close solution is disfavored by $\Delta \chi^2 = 85$, and we adopt the wide solution.

In addition, since the event shows no evidence for a caustic crossing, we have checked if the event could be explained by a single lens and a binary source. The magnification of each source is defined by the standard “Paczyński” parameters, but with a single time scale for both sources, $t_E$. 

In addition, since the event shows no evidence for a caustic crossing, we have checked if the event could be explained by a single lens and a binary source. The magnification of each source is defined by the standard “Paczyński” parameters, but with a single time scale for both sources, $t_E$. 

Figure 3. The recovered probability distribution for the finite source size, setting an upper limit of $\rho < 0.007$, and showing that the data are consistent with a point source. The distribution is used to construct a prior for our Bayesian analysis.
For each dataset we introduce an additional flux coefficient,

\[ F(t) = f_{a1,1} \cdot A_1(t) + f_{a2,1} \cdot A_2(t) + f_{b,1}. \]  

We find that such a model fits the data as well as the binary-lens model (formally slightly better, \( \Delta \chi^2 = 5 \), for 1703 degrees of freedom). Gaudi (1998) has noted that a binary-source model and a planetary model can be distinguished by the color difference expected for two sources of unequal luminosity. Recently, [Hwang et al. 2013] have used this method to resolve this degeneracy for microlensing event MOA-2012-BLG-486, and found that it was a binary-source event. For our event, the apparent unlensed magnitudes of two sources would be 20.06 ± 0.08 and 21.86 ± 0.07, corresponding to G- and K-type main sequence stars, respectively, at the distance of the bulge. For such stars, this would predict a MOA–OGLE color difference between the two sources, \(-2.5\log_{10} \left[ (f_{s,MOA}/f_{s,OGLE})/(f_{s,MOA}/f_{s,OGLE}) \right] \), of ~0.2. This is inconsistent with the recovered color difference from the binary source model, of only ~0.03±0.04 mag, and rules out the binary-source model.

In many anomalous microlensing events (e.g. Udalski et al. 2003; Gaudi et al. 2008; Muraki et al. 2011; Han et al. 2013; Kains et al. 2013), high-order effects, such as microlens parallax and finite-source effects, can break (or partially break) the degeneracies among the physical parameters that determine \( t_E \). However, since the event duration was short and with moderate magnification, the amplitude of the microlens parallax effect is small, and including it does not improve the fit significantly (\( \Delta \chi^2 = 7 \)). Our modeling sets an upper limit on the Einstein-radius-normalized source radius of \( \rho < 0.007 \) (3σ level). The best-fit model has \( \rho = 0.002 \), but the data are also consistent with a point source at the \( \sim 1 \sigma \) level (\( \Delta \chi^2 = 1.4 \)). The probability distribution for the finite-source size, recovered from the MCMC chain, is shown in Figure 3.

To try to set further constraints on the source and lens properties, we construct a color-magnitude diagram (CMD) of objects within 90′ of the event’s position (Figure 4), using OGLE-III calibrated V-band and I-band magnitudes of images, taken before the event (there were no V-band images during the event). We estimate the position of the “red clump” at \((V−I, I)\) = 19.4, 15.45 and compare it to the unreddened values derived by Natal et al. (2013) for the Galactic coordinates of the event, \((l,b)=(3.6,−2.8)\). Natal et al. (2013) find \((V−I, I, d) = (1.06, 14.36)\), i.e. a line-of-sight I-band extinction to the red clump of \( A_I = 1.09 \), and a distance modulus of 14.48 ± 0.24 mag, or 7.9 ± 0.9 kpc.

Due to the strong blending near the lens position, the source star is not detectable in images before the event, and thus its color is unknown. However, the inferred source I-mag from the microlensing model, using the base flux of the blend star and accounting for the blending fraction, is 19.83±0.05. Assuming the source star is at the same distance as the red clump, its absolute magnitude is \( M_I = 4.26±0.25 \), i.e. likely a G-type main-sequence star of angular size \( \theta_a \approx 0.6 \mu as \) (Dotter et al. 2008). Combining this with our lens model’s upper limit on \( \rho < 0.007 \), a lower limit on the Einstein angle is

\[ \theta_E > 0.085 \text{ mas}. \]  

For the best fit value of \( \rho = 0.002 \), the Einstein angle is \( \theta_E \approx 0.3 \) mas. The corresponding proper motion is \( \mu = \theta_E/t_E \approx 4.7 \text{ mas/year}. \)

In principle, some or all of the bright “blend light” in this event could be due to the lens star itself. Its isolated position on the CMD, at \((V−I, I) = (1.45, 16.7)\) (see Figure 4), \( \sim 1 \) mag brighter than the general track of the stars in this region of the diagram, suggests that it might be a foreground disk star that suffers from relatively less extinction. On the other hand, our Bayesian analysis, below, suggests such a nearby lens is highly unlikely, and therefore the blend may be due to one or more unrelated stars seen in projection. In addition, based on long-term OGLE data, the proper motion of the blend star relative to the majority bulge stars in the field is very small, 2±1 mas/year, thus also suggesting it is located in the bulge. If so, it might be a horizontal branch star. Another possibility is that this blend is not the lens itself, but a companion to the lens.

4 PHYSICAL PARAMETERS - BAYESIAN ANALYSIS

The physical parameters of the lens and its companion are connected to three of the model parameters: \( t_E \) – which involves the lens mass and distance, and \( q \) and \( s \) which, given the lens properties, give the companion mass and projected distance to the lens star, respectively. Given the limited constraints we can set, in this case, using high-order effects and
the CMD, we use Bayesian analysis to estimate probabilistically the lens distance and mass, and the Einstein radius of the event, following previous analyses of microlensing events (e.g. Batista et al. 2011; Yee et al. 2012). The prior distribution includes a Galactic stellar structure model, which sets the rate equation for lensing events. As noted, due to the large amount of blended light, likely unrelated to the lens, we cannot place strong constraints on the lens mass from the observed flux.

The event rate is

$$d^3\Gamma = \frac{q^2}{dD_L dD_S dM d^2\mu}$$

$$= (2R_E V_{\text{rel}} n_L(x, y, z) n_S(x, y, z) D_S^2 f(\mu) g(M),$$

where $R_E$ is the Einstein radius, and $V_{\text{rel}}$ is the lens transverse velocity relative to the source-observer line of sight. $V_{\text{rel}}$ is related to $\mu$, the lens-source relative proper motion, by $V_{\text{rel}} = \mu D_L$. The local densities of lenses and sources are $n_L(x, y, z)$ and $n_S(x, y, z)$, and $f(\mu)$ is the lens-source relative proper motion probability distribution. Finally, $g(M)$ is the lens mass function. We incorporate our constraints on the Einstein angle, $\theta_E$ (Eq. 5), by using the MCMC probability distribution for $\rho$, to construct the prior distribution for $\theta_E$, which is assigned to each combination of $M$, $D_L$, and $D_S$ in the rate equation.

We adopt the Galactic model of Han & Gould (1993, 2002), which reproduces well the observed statistical distribution of properties of microlensing events. The stellar density, $n(x, y, z)$, includes a cylindrically symmetric disk, a bulge and a central bar (for specific model parameters see table 2 in Batista et al. 2011).

The proper motion probability distribution, $f(\mu)$, is a two dimensional (Galactic North and East directions) Gaussian distribution, with expectation value

$$\mu_{\text{exp}} = \frac{V_L - (V_\odot + V_{\text{hel}})}{D_L} - \frac{V_S - (V_\odot + V_{\text{hel}})}{D_S}.$$  

(8)

The Sun’s velocity, consisting of a random and a circular velocity component, is $(V_\odot, V_{\text{hel}}, V_{\text{disk}}, V_{\text{bulge}}) = (7, 12) + (0, 230)$ km s$^{-1}$. The Earth’s velocity, as seen by the Sun during the event, was $(V_\odot, V_{\text{hel}}, V_{\text{disk}}, V_{\text{bulge}}) = (-19.8, 9.2)$ km s$^{-1}$. $V_L$ and $V_S$ are the expectation values for the lens and source velocities, which differ for disk and bulge lenses. We adopt means and standard deviations for the disk velocities of $(V_{\text{disk, gal}}, V_{\text{disk, bulge}}) = (0, 220)$ km s$^{-1}$, and $(\sigma_{\text{disk, gal}}, \sigma_{\text{disk, bulge}}) = (20, 30)$ km s$^{-1}$, respectively.

For the bulge, $(V_{\text{bulge, gal}}, V_{\text{bulge, bulge}}) = (0, 0)$ km s$^{-1}$ with $(\sigma_{\text{bulge, gal}}, \sigma_{\text{bulge, bulge}}) = (100, 100)$ km s$^{-1}$.

For the mass function, we follow Dominik (2006) and use different mass functions for the Galactic disk and bulge, consisting of power laws and a log-normal distributions in $(M/M_\odot)$ adopted from Chabrier (2003).

The posterior probability distributions for the lens mass and distance are found by marginalizing over all other parameters, and for the Einstein radius by summing the probability for the appropriate combination of $D_L$, $D_S$ and $M$, marginalizing over $\mu$. Figure 5 shows the results of our Bayesian analysis. The inferred lens mass is $M = 0.39^{+0.45}_{-0.19} M_\odot$, and thus the companion is a planet with mass of $11.6^{+7.4}_{-3.6} M_J$. The uncertainties are the 68% probability range about the median of the probability distribution, which we take as the most likely value. The system’s distance is $7.56 \pm 0.91$ kpc. We find that the Einstein radius of the lens star is $2.38^{+0.81}_{-0.65}$ AU, so the projected separation of the companion, $r_\perp = s \cdot R_E$, is $4.3^{+1.5}_{-1.2}$ AU, rather far beyond the location of the snowline at $R_{\text{SL}} = 1.1^{+1.2}_{-0.6}$ AU [assuming a relation $R_{\text{SL}} = 2.7(M/M_\odot)$ AU].

5 DISCUSSION

We have presented the detection, via microlensing, of a Jovian planet orbiting around a likely M-type star. There are three additional microlensing events that have detected Jupiters around M stars (Udalski et al. 2003; Batista et al. 2011; Poleski et al. 2013), and together they constitute ~1/5 of the planetary systems detected to date through microlensing. In all of them, the planet is located beyond the snowline, at distances $\lesssim 5$ AU from the host star. However, within the uncertainties, for two out of the four events (Poleski et al. 2013 and this work), the primary could also be a K- or G-type star. If snowline-region massive planets around M stars are indeed common, this may be in conflict with the two leading planetary formation scenarios. According to the “core accretion” scenario (e.g. Ida & Lin 2004), Jovian planets form beyond the snowlines of their parent stars, but massive planets around M-type stars should be rare (Laughlin et al. 2004), since their formation times are longer than the typical disk lifetime. In the disk instability planet formation scenario (e.g. Boss 2002) massive planets do form around M stars, but at distances $\gtrsim 7$ AU.

Future high resolution imaging could confirm our re-
sults by isolating the light from the lens. In addition, a set of such images, spread over several years after the event, could measure the relative proper motion between the lens and the source star, and set stronger constraint on the system parameters. Since this is a general problem for many of the planets detected by microlensing, a dedicated observing program following all microlensing planets is essential for the interpretation of those planets, and could give better priors for future Bayesian analysis of such events.

Unlike most of the microlensing-detected planets to date, the planet presented here was not detected in real time, but in a post-season analysis, illustrating the essence and elegance of the second-generation survey principle. The other planetary events that were inside the collaboration footprint in the 2011 season could have also been fully characterized by the survey data alone. The final results of this controlled experiment can thus eventually help us determine the frequency of snowline planets in the Galaxy.

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