Experimental Investigation and Multi-Response Optimization of Machinability of AA5005H34 Using Composite Desirability Coupled with PCA

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Abstract: Minimum quantity lubricant (MQL) is an advanced technique in machining to achieve sustainability, productivity, higher precision, economic benefits, and a reduction in carbon footprints. The present research work aims to investigate the effect of the cutting process parameters of the end milling of AA5005H34 material under dry and MQL cutting environments. The key performance indicators of machining include the surface roughness profile, the material removal rate, and tool wear. Surface roughness parameters are measured with the help of the Mitutoyo surface roughness tester, and the cutting tool wear is measured according to the ISO 8688-2:1989 standard using a scanning electron microscope (SEM). Sixteen experiments are designed based on the Taguchi orthogonal array mixture design. Single responses are optimized based on signal to noise ratios, while for multi-response optimization composite desirability function coupled with principal component analysis is applied. Analysis of variance (ANOVA) results revealed that the feed rate followed by spindle speed, axial depth of the cut, width of the cut, and cutting environment are the most significant factors contributing to the surface roughness profile, material removal rate, and tool wear. The optimized parameters are obtained as cutting speed of 3000 rev/min, feed rate of 350 mm/min, axial depth of cut of 2 mm, and width of cut of 6 mm under an MQL environment.

Keywords: milling; minimum lubricant quantity (MQL); aluminum alloy; Taguchi orthogonal array; signal to noise ratios; analysis of variance; principal component analysis (PCA); composite desirability method

1. Introduction

Green manufacturing is a leading trend in sustainable production to produce machined jobs with the aim of substituting environment polluting processes. The objective can be best achieved by adopting dry machining [1]. Dry machining is a sustainable method that diminishes the problems associated with the machining of commercially used alloys such as steel, aluminum, and cast iron [2]. Owing to the demands of future machining, such as cost, health concerns, and mitigating ecological issues, dry machining
is an optimal choice to resolve these issues [3]. It is widely employed as a reference standard to measure cutting fluids’ effectiveness and the interrelationships among tool wear, product quality, and temperature [4]. However, machining aluminum under dry conditions is extremely difficult due to the higher chemical affinity [5]. This can be controlled by numerous cooling techniques such as flooded coolant [6], oil-water [7], cryogenic cooling (carbon dioxide, liquid nitrogen) [8], minimum quantity lubricant (MQL) [6], and nanofluids [9]. MQL is a novel cooling technique in which coolant is sprayed at the tool–surface interface with the aid of compressed air at a controlled pressure [10]. MQL reduces the total manufacturing cost as well helping to obtain vital and desirous machining quality attributes, such as excellent surface finish i.e., average roughness (Ra), root mean square (Rq), depth of peaks and valleys on the surface (Rz) [11], greater tool life [12], higher material removal rate [13], and enhanced dimensional accuracy [10]. In the MQL strategy, nearly 70% of heat is carried away by chips [14]. The MQL method is the most suitable, as it successfully satisfies the pre-requisites of sustainable eco-friendly machining [15].

In milling, the material removal rate (MRR) has vital importance, as it significantly affects the production rate/productivity, energy consumption, cutting forces, and tool life [16]. Surface roughness (SR) is an important criterion and a crucial indicator in achieving the desired quality and performance of the machined product. Hence, SR must be controlled during machining to obtain the lowest possible values [17]. Abhang et al. [18] demonstrated that SR governs the fatigue strength, friction properties, wear, and corrosion resistance of machined surfaces. Tool life is the actual service length for which a tool performs satisfactorily, provides a good surface finish, and provides a reasonable MRR before replacement. A longer tool life reduces cost and improves productivity [19]. The appropriate selection of critical parameters, especially cutting speed, feed rate, depth of cut, and lubrication method, significantly impacts the job quality and process efficiency [20]. Vegetable oil [6,21], esters, and nanofluids [6,9,21] have substituted poisonous petroleum-based oils due to their eco-friendly, biodegradable, non-toxic, and renewable nature.

Jebaraj et al. [22] studied the influence of the process parameters—namely feed rate, spindle speed, and cooling environment (wet and cryogenic CO2 and LN2)—in the milling of Al6082-T6 alloy using an uncoated carbide insert. The output responses were average surface roughness, feed force, normal force, axial force, and cutting temperature. TOPSIS (The technique for order of preference by similarity to ideal solution) was applied to optimize the multi-responses. The analysis of variance (ANOVA) results showed that cooling environment has a significant effect on all considered responses, followed by feed rate and spindle speed. The optimal parameters obtained were wet cooling conditions, cutting speed of 125 m/min, and feed rate of 0.02 mm/tooth. Raju et al. [23] developed a model for surface roughness and optimized the cutting parameters of aluminum alloy 6061 with carbide and high-speed steel (HSS) in end milling under dry and wet environments. The ANOVA results showed that for both dry and wet environments, feed rate followed by spindle speed has a dominant effect on surface roughness; however, the depth of cut was found to be the least contributing factor. A genetic algorithm (GA) was employed to achieve optimum machining conditions. For the HSS tool without the use of a coolant, the optimized parameters are a cutting speed of 1488 rpm, a feed rate of 200 m/min, and a depth of cut of 1.5 mm; however, with the use of a coolant, they are 1465 rpm, 200 m/min, and 1.5 mm. For a coated carbide tool without the use of a coolant, they are 1500 rpm, 200 m/min, and 1.5 mm, while for with the use of a coolant they are 1477 rpm, 200 m/min, and 1.5 mm.

Elsen et al. [24] optimized the end milling process parameter (spindle speed, feed rate, and depth of cut) for a stir-casted alumina-reinforced aluminum metal matrix composite using response surface methodology (RSM) under a dry environment. The performance parameters investigated were average surface roughness and machining time. The ANOVA results revealed that cutting speed and depth of cut have a significant effect on surface roughness, while for machining time the cutting speed and feed rate were the most
influencing parameters. Minimum surface roughness (1.514 μm) and machining time (18.4 s) were obtained at optimized parameters—i.e., higher cutting speed of 1750 rev/min, lower feed rate of 0.4 mm/rev, and depth of cut at 0.2 mm.

Kumar et al. [25] studied the high-speed milling of BSL 168 aluminum composites under wet conditions. The Box Behnken array is designed for experimental runs comprised of four factors—i.e., spindle speed, depth of cut, feed rate, and coolant. The output responses were surface roughness and material removal rate, while the cutting operation was performed on the flat end mill cutter. Responses are simultaneously optimized based on composite desirability function. ANOVA results revealed that the depth of cut has a significant effect on the surface roughness and material removal rate, followed by spindle speed, feed rate, and coolant. Tsao [26] applied the Grey–Taguchi method to optimize the end milling parameters of aluminum alloy A6061P-T651. The factors considered were coating type, helix angle, primary relief angle, cutter diameter, depth of cutting, width of cut, feed rate, and spindle speed. The type of cutter used was an end mill tungsten carbide and the responses were flank wear and surface roughness. The ANOVA results showed that the most influencing parameters affecting flank wear were cutter diameter followed by coating type, helix angle, relief angle, depth of cutting, and width of cutting. For surface roughness, feed rate followed by cutter diameter and coating were found to be significant.

Lmalghan et al. [27] optimized the face milling parameters of aluminum alloy AA6061 using the Response Surface Methodology and the Particle Swarm Optimization technique under dry conditions. Experimental runs are planned based on the central composite design (i.e., 20). The ANOVA results showed that the spindle speed was the most dominant factor affecting all responses, followed by feed rate and depth of cut. The optimized parameters obtained based on PSO to achieve a minimum surface roughness (0.494 μm), cutting force (166.85 N), and power consumption (0.265 kW) were spindle speed of 3000 rpm, feed rate of 500 mm/min, and depth of cut of 3 mm. Rajeswari and Amirthagadeswaran [28] investigated the machinability characteristics and optimization of the end milling of aluminum composites (7075 metal matrix) using RSM-based grey relational analysis. The machining performance to be studied were surface roughness, material removal rate, tool life, and power consumption. It was found that the spindle speed and weight percentage of SiC were the most significant parameters for all responses. The optimized parameters to achieve a higher material removal rate, low cutting force, minimum surface roughness, and lower tool wear were spindle speed of 1000 rpm, feed rate of 0.03 mm/rev, depth of cut of 1 mm, and 5% of SiC. Tosun and Pihtili [29] optimized the cutting process parameters (feed rate, spindle speed, tool material, and cooling technique) of the face milling of 7075 aluminum alloy using grey relational analysis (GRA). The responses considered were surface roughness and material removal rate. They concluded that feed rate is one of the dominant factors among the considered factors that affect the surface roughness and material removal rate in milling, followed by cutting speed, tool material, and cooling technique. The optimized parameters obtained are the cooling environment of a flood coolant, the tool material of a carbide tool, a spindle speed of 1330 rpm, and a feed rate of 80 mm/min.

The industry’s desire to improve quality, productivity, efficiency, and cost reduction is strongly associated with the proper selection of cutting parameters at optimal levels [29]. The utmost objective of the machining sector is a higher MRR, admirable quality surface finish, higher tool life, and cost reduction. However, achieving this objective is quite difficult, as finding the optimal combination that simultaneously improves surface finish, MRR, and diminishing tool wear rate requires the application of complex statistical methods and resource consumption. High feed rate, cutting speed, and depth of cut promote a higher MRR; however, this combination also causes intense heat production at the cutting zone that significantly deteriorates the dimensional accuracy, tool life, and surface finish [21,29]. Abas et al. [30] carried out experimental and statistical studies to optimize the cutting parameters to minimize cutting forces and shape deviations in Al6026-T9. The machining was performed with virgin olive oil and an uncoated tungsten carbide tool.
The experimental runs were planned based on Taguchi mixture orthogonal array design L16. Their study concluded that feed rate was the most important factor for all components of cutting forces and shape deviation. Waseem et al. [31] applied multi-response optimization to predict the tensile creep behavior of parts produced by means of additive manufacturing. The work was accomplished using a categorical response surface methodology. Their research work revealed that the most influencing factors for creep rate were layer height, infill percentage, and infill patterns. In statistical evaluation and prediction, one of the promising techniques is structural equation modeling (SEM). For example, Kamal et al. [32] used SEM to perform a quantitative analysis of the sustainable use of construction materials. Their results revealed that the components of supply chain integration were statistically significant in terms of the construction industry performance.

Besides aluminum alloys, researchers have also explored the optimum surface roughness parameters of hard-to-cut materials under different cutting conditions. For instance, Markopoulos et al. [33] investigated the slot milling of hardened AISI O1 steel alloy and evaluated the influence of cutting feed, speed, and depth of cut on the surface roughness. Their work considered the sustainability assessment criterion for determining stable characteristics for the eco-benign machining of hardened steel. Pimenov et al. [34] studied the surface quality and energy consumption of AISI 1045 steel during face milling in terms of costs and material removal rate. Abbas et al. [35] reported a detailed study on optimizing the milling process parameters of high-strength grade-H steel by employing artificial neural networks and the Edgeworth–Pareto method. Muhammad et al. [36] performed experimental tests and statistical evaluations for the micro-milling of Inconel 718 to evaluate the effect of the tool coating and cutting parameters on surface roughness. The ANOVA results predicted that the coating was a significant parameter for both burr formation and surface roughness. Sen et al. [37] investigated the wire electrical discharge machining (WEDM) process parameters of an Inconel-800 superalloy. They applied a trapezoidal interval type-2 fuzzy number integrated analytical hierarchy process-based additive ratio assessment method. They evaluated the obtained results with existing multi-criteria decision-making methods.

A local manufacturing sector is facing issues in controlling and optimizing machining parameters to achieve the desired surface finish quality, higher productivity (in terms of material removal rate), and low setup cost (in terms of low tool wear). Currently, the setting of cutting parameters is attained by commonly employed traditional methods—i.e., trial and error—which is cumbersome, error-based, and time-consuming. One of the most reliable approaches to determine the optimum cutting parameters is numerical simulation. Numerous reliable studies have shown that numerically predicted outcomes closely conform with the experimental results. For instance, Waqas et al. [38] performed a numerical modeling and analysis of Ti6Al4V alloy using Abaqus/Explicit and predicted the cutting reaction forces and shearing zone temperature. The predicted results were close to the experimental ones. A general summary of the related published studies is presented in Table 1. The summary outlines the cutting conditions used in any particular machining process, the research approaches adopted, and the machining characteristics studied.
This article aims to identify, evaluate, and optimize the end milling cutting parameters on quality responses for AA5005H34 by employing a hybrid approach of principal component analysis (PCA) coupled with the composite desirability function (CDF). Sixteen experiments were conducted based on the L₁₆ Taguchi orthogonal array (OA) mixture design to assess the optimal cutting parameter combination for quality responses—namely, surface roughness profile, material removal rate, and flank wear (Vₜ). It is reported regarding the authors’ best knowledge and literature survey that the optimization of the end milling machining of AA5005H34 has not been reported yet. Therefore, this research represents a worthwhile contribution in fulfilling the said gap.

Initially, this research presents materials and methods for selecting machining parameters. This is followed by the Taguchi orthogonal array (OA) mixture design. Then, the statistical analysis and optimization of experimental results are analyzed by the
Taguchi SN ratios. In the last section, the achievement of multi-optimization and its validation through confirmatory experiments is explained.

Applications of Research Work

It is believed that this research is the first reported study on the optimization of the end milling machining of AA5005H34. The authors could not find published articles on this study in renowned databases by applying different search strings and terms with different Boolean operators. Therefore, this study would bridge the gap in this research domain in the form of a worthwhile contribution. This research is beneficial for researchers in advanced machining as it employs MQL cutting conditions. The presented study proves that MQL produces promising results due to it offering a higher productivity, improved surface texture, and extended tool life as compared to dry machining. The experimental and statistically grounded research findings provide a comprehensive guide to the manufacturing sector for improving the machinability of AA5005-H34.

2. Materials and Methods

2.1. Material

Aluminum alloy 5005 H34 (BOZHONG METAL GROUP, Shanghai, China) was machined using a computer numerical control (CNC) end milling machine, LG-500 HARTFORD (Hartford machining centers, Shanghai, China). These alloys are widely used because of their excellent atmospheric corrosion resistance capabilities, good strength, and good weldability. The most common applications are in high strength foil, architectural applications, and general sheet metal work [48]. Its chemical compositions are shown in Table 2.

Table 2. Chemical composition of AA5005-H34.

| Element | Al | Si | Mg | Fe | Cr | Cu | Mn | Zn | Others |
|---------|----|----|----|----|----|----|----|----|--------|
| % by weight | 97.63 | 0.30 | 1.10 | 0.07 | 0.10 | 0.20 | 0.20 | 0.25 | 0.15 |

2.2. Cutting Process Parameters

The selection of input cutting process parameters and levels was carried out through a literature survey, manufacturer recommendations, machining experts, and machine capabilities. Six influential parameters are taken in this research work—i.e., spindle speed (S), feed rate (F), depth of cut (D), width of cut (W), and cutting conditions (namely, dry and MQL).

Figure 1 shows the experimental setup, illustrating the machining of aluminum alloy and also the cutting kinematics with the MQL setup. The selected parameter units and levels are depicted in Table 3. The lubricant used was olive oil. It was proven to be effective in the end milling of AISI 304 steel [49], the turning of 40 HRC hardened steel [11], the turning of AISI 1060 steel [50], and the turning of aluminum Al 6026-T9 [37]. Olive oil is also preferred because it is biodegradable and has environmentally friendly [11] characteristics. Its viscosity is 84 cp at 20 °C, its specific heat capacity is 1.97 J/(g °C), its specific gravity is 0.911 at 20 °C, and its boiling point is 700 °C.
Figure 1. Experimental setup: (a) end milling of aluminum alloy, (b) schematic of cutting kinematics and MQL setup.

| Table 3. Selected machining parameters and levels. |
|-----------------------------------------------|
| Factors            | Symbols | Units     | Levels               |
|--------------------|---------|-----------|----------------------|
| Spindle speed      | S       | rev/min   | 1000 2000 3000 4000 |
| Feed rate          | F       | mm/min    | 250 350 450 550     |
| Axial depth of cut | D       | mm        | 1 1.5 2 2.5         |
| Width of cut       | W       | mm        | 2 6 10 14           |
| Cutting conditions | C       | -         | Dry MQL - -         |

2.3. Responses Variables

The predominant quality characteristics that are considered in this research work are the surface roughness profile, the material removal rate, and tool wear. Surface roughness portrays the geometry and texture of the machined surface and greatly influences manufacturing costs. The three main surface roughness parameters depicted in standards are average roughness height (Ra), root mean square height (Rq), and average of maximum roughness (Rz) [39]. They were measured at three different locations and measurements were repeated twice at each location on the machined surface with the help of a Mitutoyo surface roughness tester (SJ-201, Mitutoyo Corporation, Kanagawa, Japan), as shown in Figure 2. The average values of readings are depicted in Table 4.

Figure 2. Surface roughness measurement.

The weight loss method was used for measuring the material removal rate (MRR) and it is expressed in Equation (1) [39].
\[ MRR = \frac{W_b - W_a}{\rho t_m} \]  

(1)

where \( W_b \) is the weight of the sample before machining, \( W_a \) is the weight of the sample after machining, \( \rho \) is the material density, and \( t_m \) is the machining time.

In the presented study, tool life is measured as flank wear that occurs on the relief or flank of the cutting tool [51]. The wear is measured according to the ISO 8688-2:1989 standard using a scanning electron microscope (SEM). According to this standard, the tool life is said to be diminished completely when the average flank wear exceeds 300 \( \mu \)m. Figure 3 shows a SEM image of the average flank wear for machining experiment number 1, as given in Table 4.

**Figure 3.** SEM of flank wear (experiment number 1).

### 2.4. Experimentation

Experiments were designed based on the Taguchi mixture design. The Taguchi orthogonal array (OA) mixture design (L16) with 16 experimental runs was selected and experiments were conducted on a vertical-type machining center, LG-500 HARTFORD (Hartford machining centers, Shanghai, China). The machining of each sample was performed with a new cutting tool. A 4 flute carbide flat end mill cutter was used, with a diameter and length of 16 mm and 75 mm, respectively. The overhung length of 20 mm was maintained in order to avoid the influence of vibration and chatter. Experimental data collected through the OA are presented in Table 4.

| Exp. No. | Parameters | Responses |
|----------|------------|-----------|
| S F D W C | S F D W C | Ra (\( \mu \)m) | Rq (\( \mu \)m) | Rz (\( \mu \)m) | MRR (mm³/s) | \( V_b \) (mm) |
| 1 1 1 1 1 1 | 1000 250 1 2 | Dry | 1.55 | 3.16 | 4.51 | 450 | 152 |
| 2 1 2 2 1 1 | 1000 350 1.5 6 | Dry | 1.72 | 3.45 | 4.92 | 970 | 196 |
| 3 1 3 3 2 2 | 1000 450 2 10 | MQL | 1.98 | 3.98 | 5.69 | 1420 | 222 |
| 4 1 4 4 2 2 | 1000 550 2.5 14 | MQL | 2.35 | 4.38 | 6.25 | 2050 | 248 |
| 5 2 1 2 3 2 | 2000 250 1.5 10 | MQL | 1.53 | 3.07 | 4.39 | 1085 | 168 |
| 6 2 2 1 4 2 | 2000 350 1 14 | MQL | 1.64 | 3.28 | 4.69 | 1250 | 175 |
| 7 2 3 4 1 1 | 2000 450 2.5 2 | Dry | 1.92 | 3.88 | 5.52 | 1415 | 243 |
| 8 2 4 3 2 1 | 2000 550 2 6 | Dry | 2.13 | 4.26 | 6.09 | 1650 | 262 |
| 9 3 1 3 4 1 | 3000 250 2 14 | Dry | 1.6 | 3.22 | 4.61 | 1290 | 215 |
| 10 3 2 4 3 1 | 3000 350 2.5 10 | Dry | 1.89 | 3.78 | 5.41 | 1850 | 269 |
| 11 3 3 1 2 2 | 3000 450 1 6 | MQL | 1.82 | 3.62 | 5.18 | 1515 | 242 |
| 12 3 4 2 1 2 | 3000 550 1.5 2 | MQL | 1.86 | 3.68 | 5.27 | 1802 | 261 |
| 13 4 1 4 2 2 | 4000 250 2.5 6 | MQL | 1.45 | 2.62 | 3.75 | 1482 | 232 |
| 14 4 2 3 1 2 | 4000 350 2 2 | MQL | 1.61 | 3.12 | 4.48 | 1655 | 258 |
| 15 4 3 2 4 1 | 4000 450 1.5 14 | Dry | 1.92 | 3.88 | 5.55 | 2100 | 295 |
2.5. Optimization Methodology

Taguchi signal-to-noise (SN) ratio, which is a mono-objective optimization process, is widely applied for identifying the optimal parameter settings of individual responses but is not used simultaneously [52,53]. The expression for the larger-the-better (LTB) S/N ratio is shown in Equation (2):

\[
S/N \text{ ratio}_{LTB} = \eta = (-10) \times \log_{10} \left( \frac{1}{m} \sum_{i=1}^{m} \frac{1}{x_{ij}} \right).
\]  

(2)

The expression for the smaller-the-better (STB) S/N ratio is shown in Equation (3):

\[
S/N \text{ ratio}_{STB} = \eta = (-10) \times \log_{10} \left( \frac{1}{m} \sum_{i=1}^{m} z_{ij}^2 \right),
\]  

(3)

where \( m \) is the number of replications and \( z_{ij} \) is the collected experimental data.

Machining processes involve multiple responses, and machining quality heavily depends on the simultaneous optimization of all quality responses [39]. Therefore, in most practical scenarios researchers demand the simultaneous optimization of all quality responses [54]. Thus, an attempt was made to complete the simultaneous optimization of quality responses using PCA coupled with CDF. The multi-optimization concept is shown in Figure 4.

The composite desirability function is a sophisticated multi-optimization technique that can be used for real-life problems and preserves the clashing of responses under harsh conditions [55]. The procedure begins with the calculation of the desirability function value. Depending upon the goal type of response optimization, the following are the three categories (STB and LTB, as expressed in Equations (4) and (5), respectively) [39]:

\[
d_{ir}(\min) = \begin{cases} 
0, & y_{ir} \leq \min y_{ir} \\
\left( \frac{y_{ir} - \max y_{ir}}{\min y_{ir} - \max y_{ir}} \right)^a \min y_{ir} \leq y_{i} \leq \max y_{ir} \\
1, & y_{ir} \geq \max y_{ir}
\end{cases}
\]  

(4)
where \( d_{ir} \) is the desirability value of the quality response for the \( i \)th treatment and \( k \)th quality response, \( Y_{ir} \) is the observed value of the response with the \( i \)th treatment, \( \text{max}Y_{ir} \) and \( \text{min}Y_{ir} \) are the maximum and minimum values of observed data at the \( k \)th response, \( T \) is the required target value, and \( \alpha \) is the assigned weight that varies from 0.1 to 10. This determines the desirability function distribution on the interval between the upper and lower target and limit.

When \( \alpha = 1 \), the desirability function is linear and equal weights are assigned to the target value and its lower and upper limits. When \( \alpha < 1 \), the desirability function will increase exponentially (concave down) and it imparts low importance to the target. When \( \alpha > 1 \), the desirability function will increase exponentially (concave upward). In this research, \( \alpha \) is set at 1, so that the desirability function is expected to increase linearly and the values will vary between [0,1]. If the observed value (\( Y_{ir} \)) exceeds the stated criteria (i.e., LTB), the desirability value approaches 1 or near to 1. However, if the observed value (\( Y_{ir} \)) is less than the stated criteria, then the desirability value becomes 0 and is unacceptable [56]. The overall composite desirability (D) is obtained by combining the individual desirability values of all responses [57] using Equation (6).

\[
P_i = \prod_{r=1}^{n} d_{ir}^w,
\]

where \( P_i \) is the combined composite desirability value for the \( i \)th treatment, \( d_{ir}^w \) is the desirability value of the \( r \)th response, and \( w \) is the weight assigned to each response when \( \sum_{r=1}^{n}w_r = 1 \). The \( i \)th treatment that yields a higher \( D \) value—i.e., 1 or a closer value—represents the best response, and the treatment setting is selected as an optimal combination. Whereas, when the \( i \)th treatment yields \( D = 0 \) or a closer value, it represents an unacceptable response [57]. Finally, the optimized milling parameters are calculated by averaging the composite desirability values at identified levels for each parameter.

The reliability of multi-objective optimization mainly relies on the method employed for determining the priority weights for each quality characteristic [58]. Usually, equal importance is given to the responses in simple problems. However, weights may vary for multi-criteria decision-making (MCDM), as effective decision-making requires appropriate weight assignment [59]. Unequal weights can be obtained from numerous weight determination methods, such as principal component analysis (PCA), entropy [60], and the standard deviation method [61]. PCA is a powerful method widely aimed at presenting correlated variables into uncorrelated principal components [62]. It is a multivariate statistical method for multi-optimization that aims to minimize vagueness, correlation, dimensions of information, and complexity [62]. PCA is a preferable method to entropy, GRA, and the standard deviation method, as it preserves actual unique information to the maximum possible extent. It focuses on conserving unique information by employing linear permutations [63]. Accordingly, PCA translates multi-objective optimization to mono-objective optimization without compromising the unique statistics [64]. Therefore, PCA is deployed to determine the relative weights for each response, and the weights are integrated with CDF to perform multi-objective optimization. For multi-response optimization cases, such as CDF, the respective weights of the responses are required for the computation of the multi-performance index (MPI), called \( D_i \), which are computed by conducting PCA on desirability values \( d_{ir} \) for each quality response [65]. The PCA method begins with the data normalization of \( d_{ir} \) followed by the computation of the correlation coefficient matrix (M), eigenvalues, and eigenvectors, and finishes with principal components (PCs) [62].
Step 01: Data Normalization: Data pre-processing: this involves the normalization of observed quality responses between 0 and 1. The data are linearly normalized in between zero and unity depending upon the criteria employed to evaluate the responses. STB criteria are adopted for Ra, Rq, Rz, and Vb and LTB criteria are employed for MRR using Equations (4) and (5), respectively.

Step 02: Development of the correlation coefficient matrix: Using computed composite desirability values \( d_{ir} \), a multi-response array \( G \) is formulated, which is presented in Equation (7).

\[
G = \begin{bmatrix}
m_1(1) & m_1(2)m_1(3) & \ldots & m_1(r) \\
m_2(1)m_2(2) & m_2(3) & \ldots & m_2(r) \\
\vdots & \vdots & \ddots & \vdots \\
m_i(1)m_i(2) & m_i(3) & \ldots & m_i(r)
\end{bmatrix}
\]  

(7)

The expression for the generation of a correlation coefficient matrix \( M \) using the multi-response array \( G \) is shown in Equation (8).

\[
M_{r,l} = \left( \frac{\text{Cov}(m_a(b),m_a(l))}{\sqrt{\text{Var}(m_a(r))}\sqrt{\text{Var}(m_a(l))}} \right),
\]  

(8)

where \( a = 1, 2, \ldots, i \), \( b = 1, 2, \ldots, r \), \( l = 1, 2, \ldots, r \), and \( \text{Cov}(m_a(b),m_a(l)) \) is the covariance of sequences \( m_a(b) \) and \( m_a(l) \).

Step 04: Eigenvalue and eigenvector evaluation: The eigenvalues \( \lambda_r \) and eigenvectors \( \psi_r \) are generated using the \( M_{r,l} \) matrix, as per Equation (9):

\[
(M - \lambda_r I_r) \times \psi_r = 0.
\]  

(9)

Step 05: Find principal components and weights: The uncorrelated principal components \( Z_{ar} \) are developed with Equation (10):

\[
Z_{ar} = \sum_{i=1}^{n} m_i(a) \times \psi_r.
\]  

(10)

The computed principal components (PC’s), eigenvalues, and eigenvectors are arranged with respect to their corresponding variance contribution.

Figure 5 summarizes the research methodology deployed in the present study.
3. Results and Discussion

3.1. Analysis of Variance (ANOVA)

Analysis of variance (ANOVA) was employed for analyzing the significance and percentage contribution of parameters regarding the response variables and suitability of experimental data for further analysis. ANOVA was conducted at a 95% confidence interval and we obtained results for the surface roughness parameters (i.e., Ra, Rq, Rz), MRR, and Vb. The ANOVA results are depicted in Table 5. A p-value less than or equal to 0.05 shows the significance of the parameters. It is evident from Table 5 that feed rate followed by spindle speed and axial depth of cut has a significant effect on the surface roughness profile (Ra, Rq, and Rz), material removal rate (MRR), and tool flank wear (Vb), while cutting conditions were found to be insignificant for all the considered responses. The width of cut was found to be significant only for Ra and MRR. Further, the percentage contribution of feed rate was relatively high for all the responses—i.e., Ra (68.27%), Rq (66.32%), Rz
(66.07%), MRR (44.68%), and $V_b$ (42.98%)—thus making it one of the most significant contributing factors compared to spindle speed, axial depth of cut, width of cut, and cutting conditions. These results are in line with the published literature reports, which show the feed rate to be the most influential parameter for the surface roughness profile [23,29,42], material removal rate [29], and tool wear [26,28].

Table 5. ANOVA for individual quality responses.

| Source | DF | Adj SS  | Adj MS  | F-Value | $p$-Value | % Contribution |
|--------|----|---------|---------|---------|-----------|----------------|
| Ra     |    |         |         |         |           |                |
| S      | 1  | 0.08    | 0.08    | 14.11   | 0.00      | 9.04           |
| F      | 1  | 0.58    | 0.58    | 107.30  | 0.00      | 68.74          |
| D      | 1  | 0.09    | 0.09    | 16.23   | 0.00      | 10.40          |
| W      | 1  | 0.04    | 0.04    | 7.43    | 0.02      | 4.76           |
| C      | 1  | 0.01    | 0.01    | 1.03    | 0.33      | 0.66           |
| Error  | 10 | 0.05    | 0.01    |         |           |                |
| Total  | 15 | 0.85    |         |         |           |                |
| Rq     |    |         |         |         |           |                |
| S      | 1  | 0.38    | 0.38    | 12.20   | 0.01      | 11.48          |
| F      | 1  | 2.18    | 2.18    | 70.51   | 0.00      | 66.32          |
| D      | 1  | 0.16    | 0.16    | 5.11    | 0.05      | 4.81           |
| W      | 1  | 0.13    | 0.13    | 4.18    | 0.07      | 3.93           |
| C      | 1  | 0.13    | 0.13    | 4.30    | 0.07      | 4.04           |
| Error  | 10 | 0.31    | 0.03    |         |           |                |
| Total  | 15 | 3.29    |         |         |           |                |
| Rz     |    |         |         |         |           |                |
| S      | 1  | 0.75    | 0.74    | 11.09   | 0.01      | 11.19          |
| F      | 1  | 4.40    | 4.40    | 65.47   | 0.00      | 66.07          |
| D      | 1  | 0.33    | 0.33    | 4.88    | 0.05      | 4.92           |
| W      | 1  | 0.26    | 0.26    | 3.94    | 0.08      | 3.97           |
| C      | 1  | 0.25    | 0.25    | 3.72    | 0.08      | 3.75           |
| Error  | 10 | 0.67    | 0.07    |         |           |                |
| Total  | 15 | 6.66    |         |         |           |                |
| MRR    |    |         |         |         |           |                |
| S      | 1  | 731,149 | 731,149 | 58.95   | 0.00      | 26.88          |
| F      | 1  | 1,215,245 | 1,215,245 | 97.98   | 0.00      | 44.68          |
| D      | 1  | 345,056 | 345,056 | 27.82   | 0.00      | 12.69          |
| W      | 1  | 275,186 | 275,186 | 22.19   | 0.00      | 10.12          |
| C      | 1  | 29,241  | 29,241  | 2.36    | 0.16      | 1.08           |
| Error  | 10 | 124,029 | 12,403  |         |           |                |
| Total  | 15 | 2,719,906 |         |         |           |                |
| $V_b$  |    |         |         |         |           |                |
| S      | 1  | 9159.20 | 9159.20 | 57.10   | 0.00      | 36.26          |
| F      | 1  | 10,857.80 | 10,857.80 | 67.69   | 0.00      | 42.98          |
| D      | 1  | 3001.30 | 3001.30 | 18.71   | 0.00      | 11.88          |
| W      | 1  | 39.20   | 39.20   | 0.24    | 0.63      | 0.16           |
| C      | 1  | 600.30  | 600.30  | 3.74    | 0.08      | 2.38           |
| Error  | 10 | 1604.00 | 160.40  |         |           |                |
| Total  | 15 | 25,261.80 |         |         |           |                |

DF (degree of freedom), Adj SS (adjusted sum of square), Adj MS (adjusted mean square), F-value (Fisher test value), $p$-value (probability value).

The reliability and adequacy of the ANOVA results were analyzed by analyzing residual (error) plots, as shown in Figures 6 and 7. These plots display the data visually between residual and fitted values so that the reliability of the assumptions can be checked efficiently. The assumptions need to be satisfied for adequate ANOVA—i.e., the residuals are normally distributed and have constant variance [63,66]. It is quite evident from the normality plot of the residuals (shown in Figure 6) that most data points lie close to the fitted line at a 95% confidence interval (CI) for each response plot, thus confirming that the residuals are normally distributed. Further, the normality assumption was checked by
the Anderson Darling (AD) test at a 95% CI. The AD test is an extremely powerful and widely used statistical tool [67]. The null hypothesis of the AD test claims that the data follow a normal distribution—i.e., $p > 0.05$—whereas the alternative hypothesis denies the claim stated in the null hypothesis—i.e., $p \leq 0.05$. The $p$-value of the responses shown in Figure 6 is greater than 0.05, thus confirming the normal distribution of the residuals. The versus fit plots shown in Figure 7 verify the assumption of constant variance, depicting that the data are completely randomized on both sides of the fitted line and no patterns are identified (curvilinear and uneven spreading) [66]. Thus, the constant variance assumption is also satisfied. Based upon the satisfaction of the assumptions, normality plots, and versus plots, the data were fit for further analysis and optimization [39].

**Figure 6.** Normal probability plots of responses.

**Figure 7.** Versus fits plots of responses.
3.2. Main Effects Plots of Means for Individual Responses

Figure 8 presents the main effects of cutting parameters on the surface roughness profile (i.e., Ra, Rq, Rz), MRR, and Vb. As shown in Figure 8a–c, a decreasing trend for surface roughness profile is observed with the rise in spindle speed and under MQL conditions. However, an increase in the surface roughness profile is observed with an increase in the feed rate, axial depth of cut, and width of cut. The decrease in surface roughness with an increase in cutting speed is attributed to the thermal softening of aluminum alloys, elimination of built-up edges (BUE), decrease in chip fracture, and increase in plastic behavior [20,43]. Compared to dry machining, MQL tends to reduce the thermal distortion of the machined work piece and also flushes away the machined chips [43]. The increase in surface roughness profile with an increase in feed rate and depth of cut (both axial and width) occurs due to vibration, an increase in cutting forces, the kinematic-geometric mapping of the cutting edge in the workpiece, and heat generation at the tool and work piece interface [44,45]. Figure 8d shows a remarkable rise in MRR with the increase in spindle speed, feed rate, axial depth of cut, and width of cut under MQL conditions. This may be attributed to an increase in the plastic behavior of aluminum alloy and the fact that the application of MQL reduces the formation of longer chips and facilitates the movement of the chips away from the tool’s surface, therefore improving the MRR [68,69]. A rigorous increase in flank wear is noticed with an increase in spindle speed, feed rate, axial depth of cut, and width of cut. This effect decreases under the MQL environment, which is depicted in Figure 8e. This is attributed to a rise in cutting temperature that generates high residual stresses in the machined surface layer, thermal distortion at the tooltip, and the sticking of work material to the tooltip, therefore accelerating the flank wear [46,47]. However, using MQL provides better control over the machining temperature and reduces the accumulation of material on the cutting surface [47,70].

![Main Effects Plots for Means of Ra](a)
![Main Effects Plots for Means of Rq](b)
![Main Effects Plots for Means of Rz](c)
![Main Effects Plots for Means of MRR](d)
![Main Effects Plots for Means of Vb](e)

**Figure 8.** Main effects plots for (a) Ra, (b) Rq, (c) Rz, (d) MRR, (e) Vb.

3.3. Single Objective Optimization

Taguchi-based S/N ratios were employed for the individual optimization of responses. In this research work, responses involve contradictory objective functions for single-objective optimization—i.e., the minimization of the surface roughness profile and
flank wear and the maximization of MRR. Therefore, the smaller-the-better criteria were used for the surface roughness profile and flank wear using Equation (3), whereas the larger-the-better criteria were used for MRR using Equation (2).

It is quite evident from Table 6 and the main effect plots for the SN ratios of response variables (shown in Figure 9) that the optimal settings for parameters involve significant inconsistency among the quality responses. Therefore, the application of multi-objective optimization techniques is needed to obtain an optimal setting that satisfies all the responses simultaneously. For instance, as shown in Table 6 at experimental level 13 and Figure 9a–c, the optimized process parameters to achieve the lower surface roughness profile are cutting speed at a high level; feed rate, axial depth of cut, and width of cut at a low level; and an MQL environment. However, the optimized setting for MRR is high cutting speed, feed rate, axial depth of cut, and width of cut and MQL condition, as depicted in Table 6 at experimental run 15 and Figure 9d. Conversely, lower tool wear supports a lower cutting speed, feed rate, axial depth of cut, and width of cut under an MQL environment. This is depicted in Table 6 at experimental run 1 and Figure 9e.

Table 6. S/N ratios of responses.

| Exp. No. | Ra  | Rq  | Rz  | MRR | Vb  |
|----------|-----|-----|-----|-----|-----|
| 1        | −3.81 | −9.99 | −13.08 | 53.06 | −43.64 * |
| 2        | −4.71 | −10.76 | −13.84 | 59.74 | −45.85 |
| 3        | −5.93 | −12.00 | −15.10 | 63.05 | −46.93 |
| 4        | −7.42 | −12.83 | −15.92 | 66.24 | −47.89 |
| 5        | −3.69 | −9.74 | −12.85 | 60.71 | −44.51 |
| 6        | −4.30 | −10.32 | −13.42 | 61.94 | −44.86 |
| 7        | −5.67 | −11.78 | −14.84 | 63.02 | −47.71 |
| 8        | −6.57 | −12.59 | −15.69 | 64.35 | −48.37 |
| 9        | −4.08 | −10.16 | −13.27 | 62.21 | −46.65 |
| 10       | −5.53 | −11.55 | −14.66 | 65.34 | −48.60 |
| 11       | −5.20 | −11.17 | −14.29 | 63.61 | −47.68 |
| 12       | −5.39 | −11.32 | −14.44 | 65.12 | −48.33 |
| 13       | −3.23 * | −8.37 * | −11.48 * | 63.42 | −47.31 |
| 14       | −4.14 | −9.88 | −13.03 | 64.38 | −48.23 |
| 15       | −5.67 | −11.78 | −14.89 | 66.44 * | −49.40 |
| 16       | −5.15 | −11.08 | −14.13 | 65.34 | −48.69 |

Optimum: S4F1D4W2C2 S4F1D4W2C2 S4F1D4W2C2 S4F3D2W4C1 S1F1D1W1C1

* Optimized setting for mono-optimization.
3.4. Multi-Objective Optimization

The multi-optimization begins with the computation of normalized values of SN ratios by employing the desirability function criteria as per Equations (4) and (5). The reliability of multi-objective optimization mainly relies on the method used for determining the relative importance (weights) for each quality characteristic [58]. Table 7 presents the composite desirability values.

| Exp. No. | Ra  | Rq  | Rz  | MRR | Vb   |
|---------|-----|-----|-----|-----|------|
| 1       | 0.889 | 0.693 | 0.696 | 0.000 | 1.000 |
| 2       | 0.700 | 0.528 | 0.532 | 0.315 | 0.692 |
| 3       | 0.411 | 0.227 | 0.224 | 0.588 | 0.510 |
| 4       | 0.000 | 0.000 | 0.000 | 0.970 | 0.329 |
| 5       | 0.911 | 0.744 | 0.744 | 0.385 | 0.888 |
| 6       | 0.789 | 0.625 | 0.624 | 0.485 | 0.839 |
| 7       | 0.478 | 0.284 | 0.292 | 0.585 | 0.364 |
| 8       | 0.244 | 0.068 | 0.064 | 0.727 | 0.231 |
| 9       | 0.833 | 0.659 | 0.656 | 0.509 | 0.559 |
| 10      | 0.511 | 0.341 | 0.336 | 0.848 | 0.182 |
| 11      | 0.589 | 0.432 | 0.428 | 0.645 | 0.371 |
| 12      | 0.544 | 0.398 | 0.392 | 0.819 | 0.238 |
| 13      | 1.000 | 1.000 | 1.000 | 0.625 | 0.441 |
| 14      | 0.822 | 0.716 | 0.708 | 0.730 | 0.259 |
| 15      | 0.478 | 0.284 | 0.280 | 1.000 | 0.000 |
| 16      | 0.600 | 0.455 | 0.464 | 0.848 | 0.161 |

Therefore, PCA was utilized to compute the relative weights, which were integrated with the complicated multi-optimization technique applied in research—i.e., CDF for the extraction of the optimal solution [62]. The eigenvalues and eigenvectors were calculated as per Equation (9) and the PCs computed as per Equation (10), and the results are presented in Table 8. The percentage of total variation provided by PCA was used as a priority weight value [65]. Owing to this, the eigenvalue associated with the first principal component (PC) accounts for 77% of the variation in five quality responses; therefore, the eigenvector for the first PC was used in the computation of the relative weights for the quality response [62].
Table 8. Principal component analysis.

| Component | PC 1     | PC 2     | PC 3     | PC 4     | PC 5     |
|-----------|----------|----------|----------|----------|----------|
| Eigenvalue| 3.852    | 1.027    | 0.0993   | 0.0216   | 0.0001   |
| Variation (%) | 0.77   | 0.205    | 0.02     | 0.004    | 0        |
| Cumulative (%)   | 0.77   | 0.976    | 0.996    | 1        | 1        |
| Eigenvector | 0.484   | −0.271   | −0.281   | 0.783    | −0.013   |
|              | 0.475   | −0.351   | 0.155    | −0.347   | 0.712    |
|              | 0.476   | −0.346   | 0.149    | −0.372   | −0.702   |
|              | −0.404  | −0.564   | 0.659    | 0.291    | −0.01    |
|              | 0.388   | 0.605    | 0.664    | 0.207    | −0.004   |

The eigenvectors of the first PCs were squared for the computation of the relative weights of each performance index [56]. The weights 0.2343, 0.2256, 0.2266, 0.1632, and 0.1505 were assigned to Ra, Rq, Rz, MRR, and Vb, respectively, as shown in Table 9.

Table 9. Variance contribution of quality responses for PC 1.

| Response Variable | Contribution |
|-------------------|--------------|
| Ra                | 0.2343       |
| Rq                | 0.2256       |
| Rz                | 0.2266       |
| MRR               | 0.1632       |
| Vb                | 0.1505       |

The weighted composite desirability values for quality characteristics were computed by assigning respective priority weights as exponents to the composite desirability values (as per Figure 4). Using Equation (6), the product of the weighted composite resulted in overall desirability. This is illustrated in Table 10. The optimal settings were obtained by averaging the $D_i$ values at the respective levels of each parameter, as illustrated in Table 11.

Table 10. Weighted and overall composite desirability.

| Exp. No. | Ra   | Rq   | Rz   | MRR  | Vb  | Overall Desirability ($D_i$) | Rank |
|----------|------|------|------|------|-----|-------------------------------|------|
| 1        | 0.97 | 0.92 | 0.92 | 0.00 | 1.00| 0.0000                        | 14   |
| 2        | 0.92 | 0.87 | 0.87 | 0.83 | 0.95| 0.5402                        | 6    |
| 3        | 0.81 | 0.72 | 0.71 | 0.92 | 0.90| 0.3443                        | 12   |
| 4        | 0.00 | 0.00 | 0.00 | 0.99 | 0.85| 0.0000                        | 14   |
| 5        | 0.98 | 0.94 | 0.94 | 0.86 | 0.98| 0.7196                        | 2    |
| 6        | 0.95 | 0.90 | 0.90 | 0.89 | 0.97| 0.6619                        | 3    |
| 7        | 0.84 | 0.75 | 0.76 | 0.92 | 0.86| 0.3746                        | 11   |
| 8        | 0.72 | 0.55 | 0.54 | 0.95 | 0.80| 0.1625                        | 13   |
| 9        | 0.96 | 0.91 | 0.91 | 0.90 | 0.92| 0.6513                        | 4    |
| 10       | 0.85 | 0.78 | 0.78 | 0.97 | 0.77| 0.3956                        | 10   |
| 11       | 0.88 | 0.83 | 0.83 | 0.93 | 0.86| 0.4845                        | 7    |
| 12       | 0.87 | 0.81 | 0.81 | 0.97 | 0.81| 0.4457                        | 9    |
| 13       | 1.00 | 1.00 | 1.00 | 0.93 | 0.88| 0.8187                        | 1    |
| 14       | 0.96 | 0.93 | 0.92 | 0.95 | 0.82| 0.6365                        | 5    |
| 15       | 0.84 | 0.75 | 0.75 | 1.00 | 0.00| 0.0000                        | 14   |
| 16       | 0.89 | 0.84 | 0.84 | 0.97 | 0.76| 0.4593                        | 8    |
Table 11. Response table for multi-optimization.

| Cutting Parameters       | Mean of $D_i$ | Delta | Rank |
|--------------------------|---------------|-------|------|
|                          | 1             | 2     | 3    | 4    |             |
| Cutting speed            | 0.221         | 0.480 | 0.494*| 0.478 | 0.2731      | 2    |
| Feed rate                | 0.547         | 0.558*| 0.301 | 0.269 | 0.2917      | 1    |
| Depth of cut             | 0.401         | 0.426 | 0.448*| 0.397 | 0.0514      | 5    |
| Width of Cut             | 0.364         | 0.501*| 0.479 | 0.328 | 0.1732      | 4    |
| Cutting conditions       | 0.329         | 0.513*| -    | -    | 0.1910      | 3    |

* Optimal levels of multi-optimization.

4. Confirmatory Experiment

The reproducibility of the results, model adequacy, and validation of the optimal combination of machining parameters identified in Table 12 were verified by conducting three confirmatory experiments. The results of the confirmatory experiments show quite significant and pronounced improvements compared to the initial conditions, as shown in Table 12.

Table 12. Comparative summary of confirmatory experiments.

| Responses       | Initial Condition PCACDF | % Improvement from Initial Condition |
|-----------------|--------------------------|--------------------------------------|
| Ra              | S2F3D4W1C1               | 1.92                                 | 1.51 | 21 |
| Rq              | 3.88                     | 3.02                                 | 4.52 | 23 |
| Rz              | 5.52                     | 1851                                 |      | 18 |
| MRR             | 1415                     | 243                                  | 198  | 30 |
| $V_b$           | 243                      |                                      |      | 22 |
| Optimal Condition | 0                      |                                      |      | 22 |

5. Conclusions

This paper represents an informative contribution to our understanding of machining parameters. Lubrication methods were investigated and optimized for the milling of AA5005-H34 by applying a Taguchi orthogonal array (OA) mixture design (L16), SN ratios, ANOVA, PCA, and CDF. The obtained results are in good agreement with the published literature reports. For example, for surface roughness profile results closely conform with that of Raju et al. [23] and Tosun et al. [29]. For the material removal rate, the results closely match those of Tosun et al. [29] and Costa et al. [42]. For tool wear, the obtained results are in line with those of Tsao et al. [26] and Rajeswari et al. [28]. The results of the statistical analysis performed on experimental values are summarized as follows:

- The PCA-CDF method found S4F3D2W2C2 to be the optimal setting for the milling machining of aluminum 5005 H34.
- In terms of % contribution, the ANOVA revealed that the feed rate is the most influential machining parameter for the SR profile. For example, the % contributions of Ra, Rq, and Rz for feed rate are 68.74, 66.32, and 66.07, respectively.
- The Taguchi S/N ratio identified that the optimal cutting parameter for the minimization of SR (Ra, Rq, Rz) was S4F1D1W1C1—i.e., a spindle speed of 4000 rev/min, a feed rate of 250 mm/min, a depth of cut of 2.5 mm, a width of cut of 6 mm, and cutting conditions using MQL. The maximum MRR was achieved at S4F3D2W4C1—i.e., spindle speed of 4000 rev/min, feed rate of 450 mm/min, depth of cut of 1.5 mm, width of cut of 14 mm, and dry cutting conditions. For the minimization of $V_b$, the optimal levels identified were S4F1D1W1C1—i.e., spindle speed of 1000 rev/min, feed rate of 250 mm/min, depth of cut of 1 mm, width of cut of 2 mm, and dry cutting conditions.
The maximum MRR obtained under dry and MQL conditions was 2100 mm³/s and 2050 mm³/s, respectively.

- The contribution of the response variables in the principal component analysis in descending order was Ra at 23.43%, Rz at 22.66%, Rq at 22.56%, MRR at 16.32%, and Vs at 15.05%, respectively.
- The optimal machining parameters identified by the integrated approach of the PCA-CDF were spindle speed of 3000 rev/min, feed rate of 350 mm/min, depth of cut of 2 mm, width of cut of 6 mm, and MQL conditions.
- Finally, confirmatory experiments based on PCA-CDF list a significant diminishing of 21% in Ra, 23% in Rq, 18% in Rz, and 22% in Vs and an improvement of 30% in MRR. Therefore, the applied technique proved to be quite effective for multi-optimization.
- The application of MQL proved to be more efficient in comparison with dry machining due to its higher productivity, improved surface texture, and extended tool life.
- The experimental statistically optimized research findings can be offered as guidelines for the manufacturing sector. Additionally, these optimal machining conditions can be used as a baseline for improving the machinability of AA5005-H34.
- For future research work, the scope of the investigation could be enhanced by incorporating additional machining parameters to study their effect on quality responses. For example, the influence of the cutting reaction force, machining-induced stress, and nano-cutting fluids could be investigated.

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Nomenclature

- **MQL**: minimum quantity lubricant
- **SEM**: scanning electron microscope
- **PCA**: principal component analysis
- **Ra**: average roughness
- **Rq**: root mean square
- **Rz**: depth of peaks on the surface
- **MRR**: material removal rate
- **SR**: surface roughness
- **RSM**: response surface methodology
- **C**: cutting condition
- **Vs**: flank wear
- **S**: spindle speed
- **F**: feed rate
- **D**: depth of cut
- **W**: width of cut
- **Wb**: weight of the sample before machining
- **We**: weight of the sample after machining
- **ρ**: material density
\( t_m \) machining time
\( SN \) Taguchi signal-to-noise (\( \text{S/N} \)) ratio
\( \text{LTB} \) larger-the-better (LTB) \( \text{S/N} \) ratio
\( \text{STB} \) smaller-the-better (STB) \( \text{S/N} \) ratio
\( M \) number of replications
\( x_{ij} \) experimental data.
\( d_i \) desirability value of quality response
\( Y_{ir} \) observed value of the response
\( y_{ir} \) observed data at the \( k \)th response
\( T \) target value
\( a \) assigned weight (varies 0.1–10)
\( D_j \) combined composite desirability
\( d_{ir}^* \) desirability value of the \( r \)th response
\( w \) assigned weight to each response
\( \text{MCDM} \) multi-criteria decision making
\( \text{MPI} \) multi-performance index (MPI)
\( M \) coefficient matrix
\( G \) multi-response array
\( \lambda_r \) eigenvalues
\( V_r \) eigenvectors
\( \text{BUE} \) built-up edges (BUE)
\( \text{PC} \) principal component

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