Examination of the factorial model of a scale developed to assess body satisfaction in the Brazilian context: a study with people 18 to 40 years old

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Received: 16 August 2020 / Accepted: 19 December 2020 © The Author(s), under exclusive licence to Springer Nature Switzerland AG part of Springer Nature 2021

Abstract
Purpose Confirmatory factor analysis was employed to investigate an instrument developed to assess body satisfaction of Brazilian women and men and to identify participants’ body satisfaction level.
Methods Brazilian young adults completed the Body Satisfaction Situational Scale and a sociodemographic questionnaire. A total of 1481 individuals (female = 1035; male = 446) aged between 18 and 40 years old participated in the study. Factorial, convergent and discriminant validity and reliability were evaluated. An invariance test was performed across sexes using multi-group analysis. The prevalence of body satisfaction among participants was calculated using the final models of the instrument.
Results The complete model of the scale (23 items and four factors) was invariant across sexes, but it did not fit the samples even after refinement. Therefore, a theoretical investigation of the scale content was performed based on literature. Thus, a reduced model composed of two factors and ten items was found for each sex. These models showed good validity and reliability to independent samples. About prevalence, most of the women were not at all satisfied with body fat and most of the men were slightly satisfied with body musculature. Further, women and men were moderately satisfied with their face, hair, and skin.
Conclusion A reduced model of the instrument for women and another for men showed adequate indices of construct validity and reliability to samples. The most participants were not very satisfied with their bodies. The results can be useful to develop protocols aiming to promote body satisfaction.
Level of evidence Level V, descriptive cross-sectional study.

Keywords Body satisfaction · Body dissatisfaction · Scale · Validity · Reliability

Introduction
Historically, researchers and clinicians have sought to investigate a variety of unusual forms of body perception and experience reported by individuals [1]. Thus, the study of body image has been extended, and currently, a large set of information is known about this subject. Grogan [2] stated that body satisfaction is the key to understanding the development of work on positive body image. Body satisfaction is the degree of appreciation that individuals have with their body or specific parts of their body (e.g., arms, legs, and abdomen). A lower body satisfaction may lead to poorer self-care in terms of health behaviors, such as vomiting, fasting as well as the use of laxatives and diet pills for weight control [3]. Therefore, a relationship exists between poor body satisfaction and increased risk of disordered eating patterns [4]. Furthermore, Cash and Smolak [1] and Grogan [2] highlight that poor body satisfaction represents probably the major factor for increasing levels of low self-esteem, anxiety, and depression.

Negative thoughts and feelings about one’s own body can be defined as characteristic of body dissatisfaction [2, 5, 6] in the attitudinal approach presented by Cash and Smolak...
Many aspects are investigated to assess body dissatisfaction, such as size, shape, muscle tone, and weight [3, 7, 8]. On the other hand, this concept also can refer to an individual’s contemptuous opinions of his/her body (i.e., perceptual approach) and has been the target of several studies aiming to identify a perceived discrepancy between actual body and ideal body [9].

Prevalence of body dissatisfaction is high in different countries, affecting more than 50% of the adult population [5, 9, 10]. This negative evaluation of the body can be an important predictor for the development of mental disorders, such as binge eating, body dysmorphia, and social anxiety disorder [5, 11]. Therefore, body dissatisfaction is linked with numerous negative consequences in both women and men, and this can be recognized as a public health problem [10].

Both body satisfaction and dissatisfaction are investigated in men and women of all ages. Body satisfaction is consistently higher in men than women [4]. On the other hand, throughout the 1990s, body dissatisfaction increased in both sexes, particularly pertaining to leanness, fat, and muscularity [1, 7, 12]. The literature [4, 5, 10] shows that women have higher levels of body dissatisfaction than men; however, both have some level of body concern. Most women want to lose weight and have a slim “hourglass” shape, whereas men want to be heavier (i.e., more muscular) and have an inverted triangle shape [2, 3]. To achieve these shapes, women and men change behaviors (e.g., eating and level of exercise). Thus, assessing an individual’s body satisfaction/dissatisfaction is relevant to identify problems and provide treatments to avoid compromising physical and mental health.

In the last few decades, scientists have extensively invested in the development of measures to evaluate body satisfaction/dissatisfaction [7, 13]. These tools encompass aspects, such as weight, shape, muscle, fat, and size [7]. The literature contains instruments to assess body satisfaction/dissatisfaction, such as Body Parts Satisfaction Scale for Men-BPSS-M [12], Body Satisfaction Scale–BSS [14], Body Image States Scale–BISS [15], and Body Satisfaction Situational Scale–ESSC [16]. BPSS-M and BSS were developed for male and female public, respectively, and this limits their use in other samples. The BISS encompass only items to assess satisfaction with the body as a whole. On the other hand, ESSC investigates satisfaction considering the body as a whole and particular parts of women and men.

Hirata and Pilati [16] proposed the ESSC in the Portuguese language for the Brazilian context. Scale items were constructed by psychometric and psychology experts that adapted eighteen items from body image scales developed for women (e.g., Eating Disorder Inventory) and men (e.g., Male Body Attitudes Scale), and they added five items considering positive and negative feelings of people about attractiveness, shape, size, muscle tone, and fat. The experts also separated the items into four factors (see Fig. 1, original model) and this structure was supported by exploratory factor analysis. With this structure, the authors performed a preliminary validation of ESSC using a Brazilian university sample, but no psychometric study was carried out a posteriori to verify the adequacy of the four-factor model in other contexts.

Thus, the ESSC is available for use in research or clinical protocols to assist in the investigation of body satisfaction/dissatisfaction in Brazilians who, according to the literature, are a group that strongly value body image [3, 9, 17]. However, we do not know if the results produced by this instrument are valid and reliable since its factorial model has never been evaluated.

In this context, estimating the psychometric properties of psychological measures (i.e., questionnaires and scales) is essential to guarantee the quality of the gathered data [18]. This assessment should be performed using robust statistical methods (e.g., confirmatory factor analysis) that are recommended to support the use of the instrument in different samples and thus produce valid and reliable data [19].

Our first objective was to estimate the psychometric properties of the original ESSC model (i.e., with four factors) for a Brazilian sample composed of the men and women in the community. Considering the differences between the sexes when assessing the body image construct, the objective of this study was to perform the psychometric assessment of the instrument separately for women and men. As a secondary aim, we determined the prevalence of body satisfaction among the participants using the fitted model of the scale.

**Methods**

**Study design and sample size**

This is an observational cross-sectional study. We calculated the minimum sample size using Monte Carlo simulation [20] and adopted the following criteria: (a) bias of parameters estimates smaller than 10%; (b) coverage of 95% confidence intervals larger than 91%; and (c) percentage of significant coefficients (power) larger or equal to 80% [21]. A total of 1000 replications employing sample sizes of 100, 150, 200, 250, and 300 were simulated. A minimum sample size of 200 was shown to be enough to attain good parameters. Therefore, this estimate was adopted for each sex, since the analyses were performed separately for women and men.

**Participants**

Brazilians of both sexes were invited to participate in the study. The inclusion criteria were literate subjects between 18 and 40 years old. The exclusion criteria were as follows:
pregnant, lactating, or puerperal women, people with total visual impairment, and individuals who reported being in treatment for any eating disorder.

To characterize the participants, a questionnaire was prepared with information about sex (female or male), age, marital status (single, married, divorced, or widowed), race (Asian, white, mixed, or black), physical activity practice (yes or no), frequency of physical activity by days of the week, aesthetic procedures performed exclusively for body change, educational level, body weight, and height. The weight and height were used to calculate the body mass index (BMI) and anthropometric nutritional status [22]. In addition, individuals completed a series of questions developed by the Brazilian Market Research Association [23] about the number of items present in the household (e.g., bathroom and personal computers) to determine the economic level of the participant.

**Measure**

The ESSC was proposed in the Brazilian context to assess body satisfaction [16]. In the original study, a set of 33 items was initially developed; however, some items were excluded, and the initial item numbering was maintained. Thus, the final version of the instrument presented in the article consisted of 23 items (nine reversed items: 2, 4, 7, 13, 16, 18, 20, 23, and 26) with 5-point Likert response scale (1 = strongly disagree to 5 = strongly agree). The factorial model of the ESSC was presented with four factors, namely F1 = “Dissatisfaction and Fat”, F2 = “External Parts”, F3 = “Satisfaction and Muscle”, and F4 = “Lower Parts”. In the original study, the authors reported acceptable values of the internal consistency for ESSC factors (alpha de Cronbach $[\alpha]=0.65–0.82$) and a good ability to discriminate groups of men and women satisfied with their body ($p<0.001$).
Procedure and ethical aspects

The recruitment process of the participants was non-probabilistic. First, the study was released by e-mail and social media (e.g., Facebook and Instagram) among students, staff, and professors of São Paulo State University (UNESP, campus Araraquara). Those individuals who are interested in the study received information about the purpose of the research. In sequence, the individuals were asked to share the research with other people (e.g., family members and peers). Thus, we used a snowball non-random sampling for data collection. Those who agreed to participate and met the inclusion criteria signed an informed consent form and completed the Portuguese paper-and-pencil version of the ESSC and the other questions to characterize the sample. Participants provided the data in a room at the university designed for the research, which accommodated at the same time five people—each in an armchair. When the participant finished filling out the questionnaire, he/she handed it to the researcher outside the room.

Ethics approval was obtained from the Human Research Ethics Committee of the School of Dentistry (UNESP, campus Araraquara), with protocol number 88600318.3.0000.5416.

Data analysis

The psychometric sensitivity of the ESSC items was evaluated by the skewness and kurtosis absolute values. The estimates lower than 3 and 7, respectively, were used as parameters of adequate psychometric sensitivity.

The construct validity of the original ESSC model (i.e., 23 items and four factors) was assessed based on the factorial, convergent, and discriminant validity. These analyses were performed separately for women and men using half of the total sample obtained for each sex.

Factorial validity was evaluated by confirmatory factor analysis (CFA) using Weighted Least Squares Mean and Variance Adjusted (WLSMV) estimator for the polychoric correlation matrix. In CFA, the goodness-of-fit indices used were chi-squared by degrees of freedom ratio ($\chi^2/df$), Comparative Fit Index (CFI), Tucker-Lewis Index (TLI), Weighted Root Mean Square Residual (WRMR), and Root Mean Square Error of Approximation (RMSEA) with 90% confidence interval. Values of the $\chi^2/df \leq 5.0$, CFI $\geq 0.90$, TLI $\geq 0.90$, WRMR $< 1.0$, and RMSEA $\leq 0.10$ were considered acceptable for model fit [18, 24, 25]. Furthermore, the standardized factorial weight ($\lambda$) of each item was evaluated and values $\geq 0.40$ were considered adequate. To fit the model, the modification indices $> 11$, calculated from Lagrange Multipliers (LM), were analyzed [18, 26].

Convergent validity was evaluated by the average variance extracted (AVE) that was calculated for each factor of the ESSC. Values of AVE $\geq 0.50$ were considered adequate [27].

Discriminant validity was evaluated by the coefficients of determination (i.e., square correlation [$r^2$] of two correlated factors) and AVE values. When values of the AVE for each pair of correlated factors were $\geq r^2$, there was discriminant validity [18, 27].

We also estimate the reliability of the ESSC factors. For that, the composite reliability (CR), the ordinal alpha coefficient ($\alpha$), and the omega coefficient ($\omega$) were calculated. Values $\geq 0.70$ in all indices indicated good reliability [18, 28].

Furthermore, we evaluated the factorial invariance of the original model across sexes. For that, multi-group analysis using the CFI difference ($\Delta$CFI) was used. The results were evaluated in levels to compare CFI values of the configural ($\lambda$), metric (thresholds), and scalar (residuals) models. The invariance was considered strong when the CFI reduction was lower than $-0.01$ among the models [29]. As the ESSC original model did not fit the data, a fitted model for each sex was investigated. Final models were confirmed in independent samples (i.e., using the other half of the total sample obtained for each sex).

After determining the model with the best fit for each sample, the mean score for each factor of the instrument was calculated separately for women and men using the percentiles (P) of the response scale as follows: $1.0 \leq 2.0$ (P ≤ 25) = not at all satisfied; $2.0 \leq 3.0$ (P25 ≤ P50) = slightly satisfied; $3.0 \leq 4.0$ (P50 ≤ P75) = moderately satisfied; $4.0 \leq 5.0$ (P > 75) = very satisfied. We also estimated the prevalence of each factor with 95% CI. It is important to clarify that the negative items were inverted before the final score was calculated.

IBM SPSS Statistics v. 22 (SPSS, an IMB Company, Chicago, IL, USA), MPLUS v. 7.2 (Muthén and Muthén, Los Angeles, CA, USA) and R v. 3.6.2 (R Core Team, 2019) were used to perform the analyses.

Results

A total of 1,481 individuals participated in the study (women = 1,035; men = 446). Their mean age was 25.2 (standard deviation [SD] = 5.9) years old for the female sample and 24.9 (SD = 5.6) years old for the male sample. The BMI mean values for women and men were 24.0 (SD = 4.6) kg/m$^2$ and 25.1 (SD = 4.8) kg/m$^2$, respectively. Table 1 presents the characteristics of the sample.

No item of the instrument showed discrepant values for skewness ($> 3$) or kurtosis ($> 7$) attesting to adequate psychometric sensitivity. Table 2 presents the psychometric indicators of the original model of the instrument (i.e.,
complete with four factors) for female and male samples and Fig. 1 presents the factorial weights.

The original model did not adequately fit the samples (see Table 1). The modification indices indicated many correlations between errors for items in both samples (ML range = 11.02–384.69). On the other hand, all factor weights of the items were ≥ 0.48. The multi-group analysis made across sexes showed invariance of the ESSC original model (ΔCFI: metric model − configural model = 0; scalar model − metric model = −0.01). To improve the model fit, some modifications were made. For the female sample, items #14 (I am satisfied with my chest) and #26 (I think I have too much cellulite) were deleted, as both had high covariance with the other items and factors. A correlation was also allowed between errors for items #8 (In general, I am satisfied with my muscle definition) and 22 (In general, I am satisfied with the size of my muscles) because they had ML = 132.35 and similar content. With this refinement, the model presented better fit (see Table 2, refined complete model for female sample 1). For the male sample,
a refinement of the model was also needed. Items #3 (I am satisfied with the thickness of my arms), #13 (If I was thinner, I would feel much better), and #18 (I am unhappy with my abdomen) showed high covariance with many items and different factors; therefore, they were excluded. A correlation between errors for items 8 and 22 (ML = 17.51) was also allowed due to similar content. After these modifications, the refined model presented better fit (see Table 2, refined complete model for male sample 1).

Tables 3 and 4 present estimates of convergent and discriminant validity and reliability of the instrument. Both the complete and refined models presented adequate convergent validity and good reliability for female and male samples; however, factor 4 did not attain good parameters in the female sample. Moreover, the discriminant validity was compromised between one pair of factors (F3 vs. F4) for both samples.

Thus, we conducted a theoretical investigation of the instrument based on social and positive psychology perspectives on human appearance [1], seeking to identify the most adequate model. This strategy aimed to build a model that is not based only on results of statistical tests, but on the underlying theory of the evaluated construct. In this way, we found specific contents for each sex. Thus, we decided to build a model for each sex. For women, we selected factors 1 (Dissatisfaction and Fat) and 2 (External Parts), and for men, we selected factors 2 and 3 (Satisfaction and Muscle). Factor 4 (Lower Parts) was not included in the male model because we did not find solid theoretical evidence to support it to assess men’s body satisfaction/dissatisfaction. The items of factor 4 seem to be more targeted at women; however, the content was already addressed in other items; therefore, this factor was not included in the female model. The constructed models were identified as reduced models, since they were composed of only two factors.

In the reduced model proposed for the female sample, errors for items #2 (I would like to change many things on my body) and #4 (I feel ashamed of my body) were highly correlated (ML = 24.19) and even after inserting a covariance, the model fit did not show a satisfactory improvement. Thus, we opted to exclude item 4 because it has content related to “shame” and not “body dissatisfaction”. After this modification, the final female model (see Fig. 1) was confirmed in an independent sample and showed a good fit of the data (see Table 2, refined reduced model for female sample 2).

In the reduced model proposed for the male sample, items #3 (I am satisfied with the thickness of my arms) and #25 (I am satisfied with my body) were excluded due to high correlations with other items and different factors (ML range = 13.00–87.33). Moreover, item 25 presented content without specificity (i.e., assessing satisfaction with the body as a whole) contributing to its exclusion. Thus, the final male

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**Table 2** Psychometric indicators for models of the instrument

| Model          | Number of items/factors | Sample 1 | Sample 2 |
|---------------|-------------------------|----------|----------|
| Complete      | 23/4                    | Female   | Male     |
|               |                         | (sample1) | (sample1) |
| Refined complete | 21/4                    | Female   | Male     |
| Refined reduced | 11/2                    | Female   | Male     |
| Complete      | 23/4                    | Male     |          |
| Refined complete | 20/4                    | Male     |          |
| Refined reduced | 12/2                    | Male     |          |

| χ²/df | RMSEA [CI 90%] | CFI     | TLI     | WRMR  | λ = factorial weight | e = errors with items with correlation | EI = excluded items | r = correlation estimates between factors |
|-------|----------------|---------|---------|-------|---------------------|--------------------------------------|-------------------|-----------------------------------------|
| 8.08  | .12 [11.12]    | .90     | .90     | 2.23  | .48–.92             |                                      |                   |                                         |
| 6.04  | .10 [10.10]    | .94     | .94     | 1.76  | .48–.92             |                                      |                   |                                         |
| 4.03  | .08 [10.08]    | .99     | .99     | 1.10  | .48–.95             |                                      |                   |                                         |
| 3.27  | .07 [10.07]    | .99     | .99     | 1.00  | .48–.95             |                                      |                   |                                         |
| 3.98  | .12 [11.12]    | .93     | .93     | 1.59  | .48–.94             |                                      |                   |                                         |
| 3.39  | .10 [10.10]    | .95     | .95     | 1.31  | .48–.95             |                                      |                   |                                         |
| 5.16  | .14 [10.14]    | .95     | .95     | 1.31  | .48–.95             |                                      |                   |                                         |
| 3.42  | .09 [10.09]    | .97     | .97     | .96   | .64–.91             |                                      |                   |                                         |

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model (see Fig. 1) was confirmed in an independent sample and showed a good fit of the data (see Table 2, refined reduced model for male sample 2).

Convergent and discriminant validity were adequate for two-factor models in both samples (see Table 3, refined reduced models). Furthermore, all values of CR, $\alpha$, and $\omega$ were adequate, attesting the reliability of the two-factor models (see Table 4, refined reduced models).

After the fit of the refined reduced models to the samples, we identified the need to modify some terms in the instrument. The term “situational” was removed in the name of the scale and the terms female (Female Body Satisfaction Scale [BSS-F]) and male (Male Body Satisfaction Scale [BSS-M]) were added. The term “situational” was excluded because the scale did not have items that assessed body satisfaction considering specific situations, such as “If I were thinner, I would feel much better wearing a bathing suit on the beach”. We also changed the name of the factors to clarify the content of the item set and to facilitate the interpretation of the score. Thus, the “Dissatisfaction and Fat” factor was renamed to “Satisfaction with Body and Fat”. The “External Parts” factor was renamed to “Satisfaction with Body External Parts”. The “Satisfaction and Muscle” factor was renamed to “Satisfaction with Body and Muscle”. Note that, before using the results of the female model (BSS-F), the responses of items #2, #7, #13, #18, #20 and #23 must be inverted (e.g., if an individual to mark “strongly agree” [5 points) for the item #2, the answer must be recoded to “strongly disagree” [1 point]). Thus, for both female and male reduced model, the higher the score, the greater the individual’s body satisfaction. In addition, we named all points of the 5-point Likert response scale (1 = strongly disagree, 2 = disagree, 3 = neither agree nor disagree, 4 = agree, 5 = strongly agree). This will better direct the participant to complete the instrument. Table 5 presents the English and Portuguese versions of the scale for each sex considering the fitted final models.
Considering the final models, the mean score for the female sample was 2.83 (SD = 1.27) for “Satisfaction with Body and Fat” factor and 3.52 (SD = 1.01) for “Satisfaction with Body External Parts” factor. The male sample registered 3.66 (SD = 1.01) for “Satisfaction with Body External Parts” factor and 3.04 (SD = 1.04) for “Satisfaction with Body and Muscle” factor. Figure 2 presents the prevalence of individuals classified according to mean scores for final models separately for women and men.

The majority of women (38.2% [CI 95% 35.2–41.0]) and men (37.7% [CI 95% 33.0–42.2]) were moderately satisfied with body external parts and this prevalence was equal for both sexes. Most women (33.9% [CI 95% 30.9–37.0]) were not at all satisfied with their body and fat and the most men (32.3% [CI 95% 27.4–36.5]) were slightly satisfied with body and muscle.

**Discussion**

This study presents data on the factorial model of an instrument (ESSC) developed to assess body satisfaction in Brazilian samples. Our results indicated that the four-factor model and twenty-three items were not the most appropriate for the samples. Refined complete models presented acceptable factorial and convergent validity, but discriminant validity was poor. Thus, a reduced model for each sex was built and these structures showed better psychometric indicators.

Considering the importance of psychometrically evaluating the ESSC—since it was never evaluated beyond the original study [16]—our first analytical step was to investigate the factorial validity of this instrument. In this analysis, we observed that the complete model did not fit data, thus revealing the poor validity. To fit the model, we used modification indices and observed the existence of many co-variances between item errors. This means that the contents of the items were highly related, which impaired the fit of the instrument. Even after refinement, the discriminant

### Table 4 Reliability estimates

| Model               | Number of items/factors | Sample               | Factor | CR    | α     | ω     |
|---------------------|-------------------------|----------------------|--------|-------|-------|-------|
| Complete            | 23/4                    | Female (sample 1)    | F1     | .94   | .94   | .73   |
|                     |                         |                      | F2     | .85   | .84   | .80   |
|                     |                         |                      | F3     | .91   | .90   | .88   |
|                     |                         |                      | F4     | .77   | .69   | .73   |
| Refined complete    | 21/4                    | Female (sample 1)    | F1     | .94   | .94   | .74   |
|                     |                         |                      | F2     | .85   | .84   | .81   |
|                     |                         |                      | F3     | .90   | .90   | .86   |
|                     |                         |                      | F4     | .75   | .65   | .71   |
| Reduced             | 11/2                    | Female (sample 1)    | F1     | .93   | .93   | .76   |
|                     |                         |                      | F2     | .84   | .84   | .80   |
| Refined reduced     | 10/2                    | Female (sample 2)    | F1-r   | .94   | .93   | .76   |
|                     |                         |                      | F2-r   | .74   | .72   | .70   |
| Complete            | 23/4                    | Male (sample 1)      | F1     | .95   | .94   | .80   |
|                     |                         |                      | F2     | .86   | .83   | .82   |
|                     |                         |                      | F3     | .94   | .93   | .93   |
|                     |                         |                      | F4     | .81   | .73   | .78   |
| Refined complete    | 20/4                    | Male (sample 1)      | F1     | .91   | .90   | .75   |
|                     |                         |                      | F2     | .86   | .83   | .82   |
|                     |                         |                      | F3     | .94   | .94   | .92   |
|                     |                         |                      | F4     | .80   | .73   | .77   |
| Reduced             | 12/2                    | Male (sample 1)      | F2     | .95   | .93   | .93   |
|                     |                         |                      | F3     | .94   | .86   | .92   |
| Refined reduced     | 10/2                    | Male (sample 2)      | F2-r   | .84   | .83   | .79   |
|                     |                         |                      | F3-r   | .90   | .89   | .88   |

Sample 1 = half of the total sample (n: female = 517, male = 223), Sample 2 = the other half of the total sample (n: female = 518, male = 223), CR = composite reliability, α = ordinal alpha coefficient, ω = omega coefficient. Factors of the original scale: F1 = Dissatisfaction and Fat, F2 = External Parts, F3 = Satisfaction and Muscle, F4 = Lower Parts. Factors of the scale with two factors: F1-r = Satisfaction with Body and Fat, F2-r = Satisfaction with Body External Parts, F3-r = Satisfaction with Body and Muscle.
validity of the models was poor. This result shows that the items of one factor were saturating in another factor, which aroused the need for a theoretical review of the instrument.

For the theoretical investigation, the literature was consulted to verify which body aspects are more important for each sex. Studies [5, 8] indicate that the ideal body types highly valued by women are thin and shapely. On the other hand, men define the ideal body as muscular, valuing the upper body (e.g., shoulders) as areas that affect their body satisfaction [12]. Furthermore, we found women and men's perception of their face as attractive, including hair, may influence overall satisfaction with the body [30, 31]. From this information, we build a reduced model for each sex. We decided to include only two factors for each model aiming to provide the clinical and academic community an instrument with greater feasibility of application, that is, that would capture body satisfaction using few items. However, even after using only two factors, some items did not contribute to the fit of the models; therefore, they were deleted. For example, item four was excluded because of its content, since people who experienced body shame do not simply feel dissatisfied with their bodies. They feel contemptuous of their body and wants to change it at any cost [1, 32, 33]. After modifications, refined reduced models showed adequate estimates of validity and reliability; therefore, they were considered the most appropriate for the data. However, it is important to emphasize that these models must be applied in other Brazilian contexts to verify the fit in other samples to support the adequacy of these structures.

In the last stage of our study, we identified the level of body satisfaction among the participants. Interestingly, both sexes were more satisfied with their hair, face, skin, and body hair. This finding may suggest that these items were not the target of participants' concerns, perhaps because they can be changed through aesthetic procedures and the use of cosmetics [34]. In other words, Brazilian men and women could value more body size and shape than face and hair. Petrie and colleagues [30] reported that American students were more satisfied with their faces than with their bodies. This result corroborates our findings.

Still, our results showed that most women and men had low body satisfaction, which corroborates the literature [3, 5, 7, 9, 10]. More specifically, our findings demonstrated that women were not at all satisfied with their amount of fat, and men were slightly satisfied with their muscles. This represents an alert for the development of dysfunctional behaviors, such as restrictive dieting, heavyweight lifting, and taking supplements to lose or gain weight. As a consequence, these behaviors can trigger mental disorders,
such as anorexia nervosa, bulimia nervosa as well as the development of anxiety and depression. Griffiths and colleagues [10] examined the relationship between body dissatisfaction and health-related quality of life in Australian women and men. They found increased levels of body dissatisfaction among participants and a significant association of this concept with poorer mental and physical health. Therefore, high levels of body dissatisfaction may threaten people’s psychological and physical well-being.

What is already known on this subject?

Researchers and clinicians have been looking for psychometrically appropriate measures to understand the mental representation that people build of their own body (i.e., body image). A body satisfaction scale developed in Brazilian context can be useful for this purpose; however, its validity has never been tested.

What this study adds?

We found that the complete version of the Brazilian measure was not valid and reliable for the data. Therefore, we fitted the instrument to samples and found a reduced model for each sex that can to help researchers and clinicians—who aim to assess individuals’ body satisfaction—to develop protocols to reduce body dissatisfaction.

Limitations

Although the present study has implications about the use of the measure to assess body satisfaction, some limitations should be highlighted. First, we used a self-reported measure, which can produce inexact data if participants do not respond with due attention; because of this limitation, we sought to use a large sample to reduce this bias. Second, our sample was mostly female and this may reflect the fact that women are more likely to participate in studies than men; however, the sample size used for the male sample was adequate to perform the analyses. Third, the non-probabilistic sample and limited age limit the generalization of our results to other population groups. Fourth, our study design was cross-sectional, which does not permit us to infer a cause-and-effect relationship. Fifth, we do not use additional measures to assess the external validity of the scale, nor an exploratory factor analysis to investigate the dimensionality of the instrument (i.e., based only on statistical grounds). Nevertheless, our results might be useful for future studies that aim to use body satisfaction scales.

Clinical implications

Developing specific measures to assess the elements of body image is relevant. However, such measures need to have good indicators of validity and reliability of the data collected. We found that the model originally proposed for the ESSC did not produce good psychometric indicators for Brazilian samples. Therefore, the use of this model in samples with characteristics similar to ours will probably produce inaccurate results. With the reduced models (BSS-F and BSS-M), researchers and clinicians might be able to assess more appropriately the body satisfaction of young women and men. However, this will only be possible if the context of application is similar to that of the sample in the present study. This evaluation is an important step for clinicians to understand and work effectively with patients who experience poorer body satisfaction. Nutritionists, psychologists, doctors, and other health professionals can work to counter negative messages related to body image aiming to encourage a healthy lifestyle that includes moderate exercise, nutritious eating, and body acceptance, because collaborative treatment involves a multidisciplinary team. Furthermore, clinicians should keep in mind that social predictors (e.g., media and peers) continually reinforce the desirability of the ideal body. An individual’s psychological knowledge can be useful in trying to reduce body dissatisfaction and promote positive body image for women and men.

Conclusion

Our study presented psychometric indicators of an instrument developed to assess body satisfaction in the Brazilian context (ESSC). The original factorial model of the instrument did not fit data and a new structure was investigated. A reduced model for women (BSS-F) and another for men (BSS-M) were found and fitted to samples. These models showed adequate indices of factorial, convergent and discriminant validity, and good reliability for samples. Furthermore, we found that most participants were not very satisfied with their bodies, with women being more dissatisfied with the amount of fat and men with their musculature. These findings may be useful for psychometric, epidemiological, and clinical literature regarding the presentation of models with good estimates of validity and reliability to investigate the population’s body satisfaction.

Acknowledgements We thank the School of Pharmaceutical Sciences of São Paulo State University (UNESP, Campus of Araraquara/SP) for institutional support. Further, we are grateful to Arthur Ragazzi, Patrícia Teixeira, and Lucas Campos for their contribution to data collection.

Author contributions Material preparation, data collection and analysis were performed by Wanderson R. da Silva. The first draft of the manuscript was written by Wanderson R. da Silva and the other authors commented on previous versions of the manuscript. All authors read and approved the final manuscript.
Funding  The study was financially supported in part by the São Paulo Research Foundation (grants #2017/20315-7, #2019/19590-9).

Availability of data and material  The datasets generated and analyzed during the current study are available from the corresponding author on reasonable request.

Code availability  Not applicable.

Compliance with ethical standards

Conflict of interest  The authors declare that they have no conflict of interest.

Appendix

See Table 5.

Table 5  Brazilian Portuguese and English versions of the Body Satisfaction Scale for each sex

| English| Portuguese |
|---|---|
| Female Body Satisfaction Scale (BSS-F) | Escala Feminina de Satisfação Corporal (ESC-F) |
| Instruction: Answer the statements below according to your opinion at the moment | Instrução: Responda as afirmações abaixo de acordo com a sua opinião neste momento |
| Response Options: 1 = strongly disagree, 2 = disagree, 3 = neither agree or disagree, 4 = agree, 5 = strongly agree | Opções de resposta: 1 = discordo totalmente, 2 = discordo, 3 = não concordo ou discordo, 4 = concordo, 5 = concordo totalmente |
| (2) I would like to change many things about my body | (2) Gostaria de mudar muitas coisas no meu corpo |
| (7) I am unhappy with my waist | (7) Estou descontente com a minha cintura |
| (13) If I was thinner, I would feel much better | (13) Se eu estivesse mais magra, me sentiria muito melhor |
| (18) I am unhappy with my belly | (18) Estou descontente com a minha barriga |
| (20) I think I have too much fat on my body | (20) Acho que tenho gordura demais no meu corpo |
| (23) I am unhappy with my body measurements | (23) Estou descontente com as minhas medidas corporais |
| (21) I am satisfied with my hair | (21) Estou satisfeita com os meus cabelos |
| (24) I am satisfied with my face | (24) Estou satisfeita com o meu rosto |
| (27) I am satisfied with my skin | (27) Estou satisfeita com a minha pele |
| (28) I am satisfied with the amount of hair I have on my body | (28) Estou satisfeita com a quantidade de pelos que tenho no corpo |
| Male Body Satisfaction Scale (BSS-M) | Escala Masculina de Satisfação Corporal (ESC-M) |
| Instruction: Answer the statements below according to your opinion | Instrução: Responda as afirmações abaixo de acordo com a sua opinião |
| Response Options: 1 = strongly disagree, 2 = disagree, 3 = neither agree or disagree, 4 = agree, 5 = strongly agree | Opções de resposta: 1 = discordo totalmente, 2 = discordo, 3 = não concordo ou discordo, 4 = concordo, 5 = concordo totalmente |
| (21) I am satisfied with my hair | (21) Estou satisfeito com os meus cabelos |
| (24) I am satisfied with my face | (24) Estou satisfeito com o meu rosto |
| (27) I am satisfied with my skin | (27) Estou satisfeito com a minha pele |
| (28) I am satisfied with the amount of hair I have on my body | (28) Estou satisfeito com a quantidade de pelos que tenho no corpo |
| (6) I like my current body weight | (6) Gosto do peso que tenho agora |
| (8) In general, I am satisfied with my muscle definition | (8) De modo geral, estou satisfeito com a minha definição muscular |
| (11) I like the width of my shoulders | (11) Gosto da largura dos meus ombros |
| (14) I am satisfied with my chest | (14) Estou satisfeito com meu peito |
| (15) I think my body is attractive | (15) Acho meu corpo atraente |
| (22) In general, I am satisfied with the size of my muscles | (22) De modo geral, estou satisfeito com o tamanho dos meus músculos |

*The cross-cultural adaptation of the scale for the English language was performed following international guidelines to guarantee idiomatic, conceptual, and semantic equivalences. Satisfaction with Body and Fat: items 2, 7, 13, 18, 20 and 23. Satisfaction with Body External Parts: items 21, 24, 27 and 28. Satisfaction with Body and Muscle: items 6, 8, 11, 14, 15 and 22. The numbering of the items on the instrument was maintained according to the proposal originally developed by Hirata and Pilati in 2010. [Fator “Satisfação com o Corpo e Gordura”: items 2, 7, 13, 18, 20 and 23. Fator “Satisfação com as partes externas do corpo”: items 21, 24, 27 and 28. Fator “Satisfação com o corpo e músculo”: items 6, 8, 11, 14, 15 and 22. A numeração dos itens do instrumento foi mantida de acordo com a proposta originalmente desenvolvida por Hirata e Pilati em 2010]
