Quality utilization of antenatal care and low birth weight: evidence from 18 demographic health surveys
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Abstract
Background: Low birthweight is a crucial factor in child mortality and morbidity and affects almost 20% of infants worldwide, mostly in low- and middle-income countries.
Aims: To assess the relationship between access to and quality of antenatal care and occurrence of low birth weight.
Methods: We analysed data from 18 demographic and health surveys, from 2005 to 2013, including 69,446 children. The main study outcome was birthweight < 2.5 kg, and access to and number of antenatal care visits were exposure variables. Moreover, antenatal care attendants and time of visit (trimester) were considered. Multiple logistic regression adjusted for sampling at primary and country level was utilized.
Results: At least 1 and ≥ 4 antenatal care consultations were both associated with decreased odds of low birth weight when compared to none and < 4 antenatal care consultations, respectively. Additional benefit stemmed from having skilled antenatal care attendants and the first antenatal care consultation during the first trimester.
Conclusions: Proper antenatal care coverage during pregnancy is beneficial for preventing low birth weight in low- and middle-income countries.

Introduction
In 2012, the World Health Assembly endorsed a comprehensive plan under Resolution 65.6 with specific global nutrition targets for 2025 (1). This policy included a 30% reduction in low birthweight (LBW) (2), corresponding to a reduction from 20 million to ~ 14 million neonates with birthweight < 2.5 kg (3) between 2012 and 2025. LBW affects almost one sixth of infants worldwide with > 95% of cases located in developing countries (3), and is recognized as 1 of the most influential factors on child mortality and morbidity. LBW increases mortality risk by 20–30 times (4), and contributes to 60–80% of all neonatal deaths worldwide (5,6). Surviving infants are at higher risk of pathological conditions such as infection immediately after birth and throughout the first year of life (7). LBW is also associated with morbidity later in life, such as psychosocial disorders (8), impaired cognitive function (9), coronary heart disease (10) and noninsulin dependent diabetes (11). Several risk factors are claimed to be associated with LBW, including maternal factors, pregnancy, multiple gestation, socioeconomic characteristics, drug treatment and body mass index (12–15).

At least 4 antenatal care (ANC) consultations, with the first preferably in the first trimester (6), has been a worldwide recommended policy for the last 2 decades. However, there is still inconclusive evidence on its impact on maternal and neonatal outcomes in developing countries. Some studies have shown that ANC improves birthweight (17,18), while others have shown a lack of evidence for the effectiveness of content, frequency and timing of visits in standard ANC programmes on maternal and child health (19).

Our research used data from demographic and health surveys (DHSs) in 18 countries and examined the association between adequate utilization of ANC and occurrence of LBW.

Methods
Study design
This was a population-based study of data from 18 DHSs between 2005 and 2013, which reported birthweight for at least 80% of births over the 5 years preceding the survey: Albania 2008/2009, Armenia 2010, Congo (Brazzaville) 2011/2012, Dominican Republic 2013, Gabon 2012, Guyana 2009, Honduras 2011/2012, Indonesia 2012, Jordan 2012, Kyrgyzstan 2012, Maldives 2009, Republic of Moldova 2005, Peru 2012, Philippines 2013, Sao Tome and Principe 2008/2009, Swaziland 2006/07, Tajikistan 2012, and Ukraine 2007. Detailed information on procedures and sampling techniques for all DHSs have been published elsewhere (20). Face-to-face interviews were carried out for a total of 213,752 women.
**Study population**

The study population consisted of all the latest singleton live births ($n = 77,809$) during the 5 years preceding the DHS in each country. After excluding 8363 (10.7%) individuals for whom we had missing data on BW, the final sample included 69,446 babies. Information on BW was obtained through birth certificates and maternal recall for 21,334 (30.7%) and 48,112 (69.3%) infants, respectively.

**Outcome, exposure and control variables**

The main outcome was LBW, which was defined as < 2.5 kg. Characteristics of ANC were the exposure variables, which were defined as follows: “ANC”, if any ANC consultation was performed; “provider”, classified as skilled (doctors, nurses or other trained attendants) or unskilled attendant (traditional attendants or others); “number of ANC consultations”, subdivided into < 4 or ≥ 4 ANC visits; “ANC timing”, dichotomized into ANC first consultation in the first trimester or after the first trimester; and “quality of ANC”, with ≥ 4 ANC consultations started in the first trimester with a skilled attendant on 1 side, and all the others on the other side. A series of socioeconomic, pregnancy and maternal characteristics were evaluated as possible confounders, including age, education, wealth, place of residence, birth interval, birth order, wanted pregnancy and child sex.

**Statistical analysis**

Statistical analysis was performed using Stata 13.1 SE (StataCorp, College Station, TX, USA). The “svy” command was used to adjust for clustering by primary sampling unit. Number of total livebirths and LBW by country were tabulated with relative percentages. All the study categorical confounding variables were tested against LBW using the $\chi^2$ test. Furthermore, we used the $\chi^2$ test to examine the association between the quality of ANC and the following socioeconomic variables: wealth status, maternal age and education, and place of residence.

The Metaprop syntax (21) was used in the pooled meta-analysis of all country datasets, which generated weighted subgroup and overall pooled estimates with inverse-variance weights obtained from a random-effects model. In this model, no residual heterogeneity was assumed. The final model included wealth, age, birth order, birth spacing, education, wanted pregnancy, child sex, and rural/urban residence; the factors primary sampling unit and country were added with random effect. Stepwise logistic regression analysis of LBW on the 5 ANC exposure variables was conducted adjusting for socioeconomic, maternal and pregnancy characteristics. $P < 0.05$ was considered statistically significant.

**Ethical approval**

This study used existing data obtained from ORC Macro (Calverton, MD, USA) through formal request mechanisms (https://dhsprogram.com). No additional ethical review for the secondary analysis was required since each country and the Institutional Review Board of ORC Macro approved the DHS data collection procedures.

**Results**

Overall, 6238 (9.0%) newborns with LBW were recorded, ranging from a minimum of 36 (2.8%) of 1281 in Albania to a maximum of 883 (20.8%) of 4238 newborns in the Philippines (Table 1). Data on ANC were missing for 1404 individuals, corresponding to 2% of the total study population. Most mothers had ANC ($n = 66,513$; 97.7%) and half of them ($n = 33,038$) had the first consultation during the first trimester (Table 2). Only 6517 (10%) women had < 4 consultations. Almost all pregnant women (97.7%) consulted a trained operator. Less than half of them ($n = 31,372$) had a good quality of ANC according to World Health Organization (WHO) criteria.

Table 3 shows a clear trend toward increasing prevalence of LBW with decreasing wealth, poorer education and shorter birth intervals, in addition to higher risk in unwanted pregnancies and female sex.

There were associations between wealth status and education and the quality of ANC. The richest and most educated women, in addition to those living in urban areas, were more likely to have ≥ 4 ANC consultations performed by skilled attendants, with the first consultation during the first trimester.

The adjusted logistic regression showed a significant benefit of having any ANC consultation when compared

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**Table 1 Numbers of live births and low birth weight infants in 18 low- and middle-income countries**

| Country, year       | Live births | Low birth weight (%) |
|---------------------|-------------|----------------------|
| Albania 2008/2009   | 1281        | 36 (2.8)             |
| Armenia 2010        | 1139        | 1,438 (6.0)          |
| Congo (Brazzaville 2011/2012) | 5355 | 467 (8.7)          |
| Dominican Republic 2013 | 2847 | 378 (13.3)          |
| Gabon 2012          | 3485        | 445 (12.8)           |
| Guyana 2009         | 1294        | 167 (12.9)           |
| Honduras 2011/2012  | 7062        | 654 (9.3)            |
| Indonesia 2012      | 13,045      | 840 (6.4)            |
| Jordan 2012         | 6612        | 817 (12.4)           |
| Kyrgyzstan 2012     | 3089        | 147 (4.8)            |
| Maldives 2009       | 3206        | 328 (10.2)           |
| Republic of Moldova 2005 | 1350 | 63 (4.7)            |
| Peru 2012           | 7385        | 476 (6.5)            |
| Philippines 2013    | 4238        | 883 (20.8)           |
| Sao Tome and Principe 2008/2009 | 1159 | 79 (6.8)            |
| Swaziland 2006/2007 | 1788        | 116 (6.5)            |
| Tajikistan 2012     | 2955        | 197 (6.7)            |
| Ukraine 2007        | 2156        | 74 (3.4)             |
| Total               | 69,446      | 6238 (9.0)           |

Results are total number of newborns and number of low birthweight infants among the last births for each woman in the preceding 5 years. Results from 18 demographic health surveys.
to no ANC (OR 1.2; 95% CI 1.0–1.4) (Table 4). Among infants of women who underwent ANC, having < 4 consultations, first consultation after the first trimester, being attended by an unskilled operator and not meeting WHO quality criteria were associated with 1.5 (95% CI 1.4–1.7), 1.1 (95% CI 1.0–1.2), 1.2 (95% CI 1.1–1.4) and 1.1 (95% CI 1.0–1.2) increased ORs of LBW, respectively.

**Discussion**

This secondary analysis of DHS data from 18 countries showed that the absence of ANC consultation increased the risk of LBW. All WHO criteria, separately and combined, for adequate antenatal consultations resulted in significant protection against LBW. We compared our results on the country incidence of LBW with other sources and found no substantial differences. Estimates from the United Nations Children’s Fund and WHO global and country reports on LBW confirm the smallest percentage (3%) for Albania up to the highest (20%) in the Philippines (22).

Our findings on the influence of maternal education on LBW are not surprising. A study in the Islamic Republic of Iran showed that the prevalence of LBW in infants born to women with no education was 16.9%, which decreased to 5.4% in women educated to a higher level (23). The explanation may lie in greater access to ANC and better nutritional behaviour. Similarly, parity and birth spacing have been detected as important determinants for LBW. One study showed that mothers with very short interpregnancy intervals (IPIs; < 3 months) and high parity had a higher risk of having LBW infants when compared to those with very short IPI but low parity (24). The explanation for these differences may be depleted nutritional reserves in women with high parity and short IPI.

Other DHSs from single countries have reported the benefit of an early start to ANC and the importance of a sufficient number of consultations. A study from Nepal showed how women with no ANC were twice as likely to have LBW infants when compared to mothers with ≥ 4 ANC consultations (25). A study from Colombia reported that having the first ANC after the first trimester was associated with an increased OR for LBW when compared with first visits at the first trimester (26). Similar findings were reported in a study in Kenya (27), indicating a positive effect of ANC, which influences dietary behaviour and treatment from any illness that may have negative effects on the health of the fetus.

Although our secondary analysis had advantages, such as large sample size and use of standardized questionnaires that limited the risk of intercountry variation, it had some limitations. First, we considered only the 18 DHSs with at least 80% of data on BW, but we cannot exclude bias for all remaining women not able to report information, which may have led to underestimation of LBW. Second, two thirds of the information on BW relied on maternal recall, therefore presenting a particular type of misreporting called heaping. Heaping consists of rounding and reporting weights as multiple of 500 g, which makes interpretation difficult when infants are reported as weighing 2.5 kg, and thus likely to be misclassified as having normal weight (28). Third, several possible confounding variables such as genetics and maternal history of diseases were not available. Finally, we had no information on nutritional status of women to exclude maternal factors that would increase risk of LBW.

### Table 2 Distribution of ANC variables among 69,446 low birthweight and normal weight livebirths in 18 low- and middle-income countries in 2005–2013

| ANC variables                  | Low birth weight | Normal weight | P     | χ²  |
|-------------------------------|-----------------|---------------|-------|-----|
| ANC                           |                 |               |       |     |
| No                            | 177 (11.6)      | 1,352 (88.4)  | < 0.001 |
| Yes                           | 5,948 (8.9)     | 60,565 (91.1) |       |
| Time of first ANC consultation |                 |               |       |     |
| First trimester               | 2,795 (8.5)     | 30,243 (91.5) | < 0.001 |
| After first trimester         | 3,248 (9.4)     | 31,406 (90.6) |       |
| No. of ANC consultations      |                 |               |       |     |
| ≥ 4                           | 5,108 (8.5)     | 54,888 (91.5) | < 0.001 |
| < 4                           | 840 (12.9)      | 5,677 (87.1)  |       |
| ANC attendant                 |                 |               |       |     |
| Skilled                       | 5,908 (8.9)     | 60,291 (91.1) | < 0.001 |
| Unskilled                     | 217 (11.8)      | 16,26 (88.2)  |       |
| ANC highest quality           |                 |               |       |     |
| Yes                           | 2,628 (8.4)     | 28,744 (91.6) | < 0.001 |
| No                            | 3,610 (9.5)     | 34,464 (90.5) |       |

ANC = antenatal care.
In conclusion, our study reinforces the need to encourage pregnant women to attend ANC to reduce LBW, with its short- and long-term consequences. Policies should in particular address access to and quality of ANC among disadvantaged socioeconomic groups, which are at higher risk of LBW. Uneducated mothers are less likely to understand health messages and to be concerned about their health and nutritional status. Poorer women are less likely to afford the cost of ANC and transportation in areas where health infrastructure is distant.

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### Table 3 Characteristics of mothers of 69,446 low birthweight and normal weight infants in 18 low- and middle-income countries in 2005–2013

| Maternal characteristics | Low birth weight | Normal weight | P | \( \chi^2 \) |
|--------------------------|------------------|---------------|---|----------|
| **Maternal age (years)** |                  |               |   |          |
| 15–19                    | 568 (13.2)       | 3745 (86.8)   | < 0.001 |          |
| 20–24                    | 1541 (9.7)       | 14305 (90.3)  |          |          |
| 25–29                    | 1538 (8.1)       | 17383 (91.9)  |          |          |
| 30–34                    | 1155 (7.7)       | 13756 (92.2)  |          |          |
| 35–39                    | 838 (8.5)        | 8996 (91.5)   |          |          |
| 40–44                    | 479 (10.4)       | 4127 (89.6)   |          |          |
| 45–49                    | 119 (11.7)       | 896 (88.3)    |          |          |
| **Birth order**          |                  |               |   |          |
| 1                        | 2227 (10.6)      | 19920 (89.9)  | < 0.001 |          |
| > 1                      | 4011 (8.5)       | 43288 (91.5)  |          |          |
| **Preceding birth interval (months)** |                  |               |   |          |
| < 18                     | 397 (12.1)       | 2871 (87.8)   | < 0.001 |          |
| 18–23                    | 438 (9.2)        | 4337 (90.8)   |          |          |
| 24–35                    | 863 (8.4)        | 9347 (91.5)   |          |          |
| > 35                     | 2313 (8.0)       | 26733 (92.0)  |          |          |
| **Place of residence**   |                  |               |   |          |
| Urban                    | 2936 (8.7)       | 30829 (91.3)  | 0.01 |          |
| Rural                    | 3302 (9.2)       | 32379 (90.7)  |          |          |
| **Education**            |                  |               |   |          |
| No education             | 248 (11.7)       | 1869 (88.3)   | < 0.001 |          |
| Primary                  | 1901 (10.1)      | 16897 (89.9)  |          |          |
| Secondary                | 3094 (8.7)       | 32338 (91.3)  |          |          |
| Higher                   | 992 (7.6)        | 12063 (92.4)  |          |          |
| **Wealth index**         |                  |               |   |          |
| Poorest                  | 1882 (11.3)      | 14780 (88.7)  | < 0.001 |          |
| Poorer                   | 1540 (9.4)       | 14748 (90.5)  |          |          |
| Middle                   | 1202 (8.5)       | 13009 (91.5)  |          |          |
| Richer                   | 960 (7.8)        | 11378 (92.2)  |          |          |
| Richest                  | 654 (6.6)        | 9293 (93.4)   |          |          |
| **Wanted pregnancy**     |                  |               |   |          |
| Wanted                   | 4077 (8.5)       | 44109 (91.5)  | < 0.001 |          |
| Not wanted               | 2156 (10.2)      | 19057 (89.8)  |          |          |
| **Child sex**            |                  |               |   |          |
| Male                     | 2964 (8.2)       | 33031 (91.8)  | < 0.001 |          |
| Female                   | 3274 (9.8)       | 30177 (90.2)  |          |          |
### Table 4 Odds ratios for low birthweight in 69 446 singleton births

| ANC                          | OR (95% CI) unadjusted | OR (95% CI) adjusteda |
|------------------------------|------------------------|-----------------------|
| No ANC visit                 | 1.3 (1.1–1.6)          | 1.2 (1.0–1.4)         |
| < 4 ANC visits               | 1.6 (1.5–1.7)          | 1.5 (1.4–1.7)         |
| ANC visit after first trimester | 1.1 (1.0–1.2)          | 1.1 (1.0–1.2)         |
| No Skilled ANC               | 1.4 (1.2–1.6)          | 1.2 (1.1–1.4)         |
| No Quality ANC               | 1.1 (1.0–1.2)          | 1.1 (1.0–1.2)         |

*aAdjusted for wealth, age, birth order, birth spacing, education, wanted pregnancy, child sex, and rural/urban residence.

ANC = antenatal care; CI = confidence interval; OR = odds ratio.

### Utilisation qualitative des soins prénatals et faible poids de naissance : données issues de 18 enquêtes démographiques sur la santé

**Résumé**

**Contexte** : Le faible poids à la naissance est un facteur crucial de la mortalité et de la morbidité infantiles et touche près de 20 % des nourrissons dans le monde, principalement dans les pays à revenu faible et intermédiaire.

**Objectifs** : La présente étude avait pour objet d'évaluer le lien entre l'accès aux soins prénataux et leur qualité d'une part, et le faible poids de naissance d'autre part.

**Méthodes** : Nous avons analysé les données de 18 enquêtes démographiques et sanitaires, de 2005 à 2013, portant sur 69 446 enfants. Le principal résultat de l'étude concernait un poids de naissance inférieur à 2,5 kg. L'accès aux visites prénatales et le nombre de consultations étaient des variables d'exposition. En outre, le personnel de consultation prénatale et le calendrier des visites (trimestrielles) ont été pris en compte. La régression logistique multiple ajustée pour l'échantillonnage aux niveaux primaire et national a été utilisée.

**Résultats** : Deux facteurs, en l'occurrence le fait d'avoir au moins une consultation prénatale et un nombre de visites supérieur ou égal à quatre, ont été associés à une diminution de la probabilité de faible poids de naissance par rapport à l'absence de visite et à un nombre de consultations prénatales inférieur à quatre, respectivement. La mise à disposition du personnel de consultation prénatale qualifié et la première consultation prénatale prévue au cours du premier trimestre de la grossesse constituaient également des avantages.

**Conclusions** : Une couverture adéquate des soins prénataux pendant la grossesse est bénéfique pour prévenir le faible poids de naissance dans les pays à revenu faible et intermédiaire.
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