Institutional factors of inclusive growth: Evidence from Côte d’Ivoire

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Economic growth is important but not sufficient to generate a sustainable increase in individual welfare. Inclusiveness attributes to growth, the dual virtue of widening the space of economic and social opportunities; while ensuring a better application of distributive equity. In this perspective, Ali and Son suggest that growth is inclusive when combined with high income and equity. This study tries to verify this assertion in Côte d’Ivoire. This study focuses on ARDL Bounds approach for testing cointegration to measure the contribution of institutional factors to inclusive growth in Côte d’Ivoire over the period of 1984 to 2018. The International Country Risk Guide (ICRG) Index is used as institutional factors. The findings of empirical analysis suggest that only government stability as institutional factors have greatly and statically significant effect on inclusive growth in the short and long run.

Key words: Ardl Bounds Test, Côte d’Ivoire, inclusive growth, institutional factors.

INTRODUCTION

The objective of strong economic growth and sustainable development in the aftermath of independence in most sub-Saharan Africa countries, particularly in Côte d’Ivoire, is to create the deployment of means and strategies to redirect economic development policy. However, in the 1960s, the economic, financial and institutional situation did not seem to be conducive to motivate economic take-off and improve the social well-being of populations.

In the late of 1980s, a compromise was reached in order to redefine the desirable macroeconomic framework for boosting growth, promoting development and reducing poverty. According to the Washington Consensus (Williamson, 1990), the institutional arrangements put in place should lead without fail to strict budgetary discipline, broadening of the tax base, privatization, deregulation, protection of private property and trade and financial liberalization; this is the conception of the “self-regulating” (Williamson,1990) and efficient market. However, such objective does not tend towards development according to Piketty, but would rather be vulnerable to Kuznets (1955) cure that economic growth feeds inequalities in the first phase of development (Asongu, 2015). Since the seminal publication of Acemoglu et al. (2005) on the crucial role of institutions quality and the economic development nations, a new literature is emerged (Bouzahzah et al., 2015; Asongu, 2015). A large number of economic studies conducted in recent year suggest that institutions are vital for economic development and growth. Indeed, when institutions have not good quality, they are influenced by power groups and act more in
their favor. As a result, significant disparities affect the redistribution of the benefits of economic dynamism to disadvantage frailest social groups, including: ethnic minorities, people living in rural areas and women of the disabled (Klasen, 2010). Since the early 1990s, debates on alternatives to the inequality literature have led to a new perception of the concept of pro-poor growth (or inclusive growth). At the same time, the concept of good institutional quality and inclusive growth is on the heart of development policy discussions and conceptions. A large number of studies have examined the role of institutions factors in economic development, poverty reduction and better, in promoting inclusive growth. Indeed, the inclusiveness of growth implies dimensions other than poverty and income distribution such as the good quality of institution factors which, in any case, has an impact on income distribution. Therefore, institutional quality and inclusive growth are important to develop policy agenda, and the question arises, whether institutional factors are conducive to the inclusiveness of Côte d’Ivoire’s economic growth.

This paper attempts to review a theoretical perspective on institutional factors of inclusive growth and estimate an empirical model to measure the contribution of institutional factors of inclusive growth, using Côte d’Ivoire as an example, where the kind research work has rarely been undertaken. The objective of this paper is to measure the contribution of institutional factors to inclusive growth. However, very few studies have considered institutional factors in explaining the inclusiveness of economic growth, especially in Côte d’Ivoire. To the best of the authors knowledge, this issue has not been the subject of any previous research for Africa and especially in Côte d’Ivoire. Even so, institutional variables have been ignored in the explanation of the results of this question. This study makes an empirical contribution for economic research by measuring inclusive growth using the method of Ali and Son (2007b) and the contribution of institutional factors to inclusive growth. Rather than being a study of the determinants of inclusive economic growth, this study contributes to the debate on the link between institutions and inclusive growth. From this point of view, this study differs from existing studies, which focus for the most part on its definition and ways of measuring inclusive growth (Klasen, 2010; Rauniyar and Kanbur, 2010). Only limited studies actually measure it and study the factors that determine it (Anand et al., 2013; Balakrishman et al., 2013; Abbe, 2019). However, these studies focus on Asia, North Africa and West African Economic and Monetary Union (WAEMU).

LITERATURE REVIEW

Indeed, most recent statistics show that sub-Saharan Africa has experienced high rates of economic growth over the past (Asongu and Le Roux, 2016). In addition, human development indexes have progressed considerably said they. While, there is general support for the notion of inclusive growth, there is no consensus. There are few empirical (or theoretical) studies on the relationship between inclusive growth and its institutional factors. The study’s analysis of the relationship between institutional factors and inclusive growth is based on the existing research and the link between economic growths. The study also gives a theoretical background in the documentation of the relationship between inclusive growth and the international country risk guide as an institutional index.

In the late 1990s, an intense debate over how institutional factors in terms of inclusive growth has been rise to divergent views and conceptions both in economic and within the international community. A large number of empirical frameworks have examined the role of quality of institution in economic development, reducing poverty and promoting inclusive growth. The necessity to improve the quality of institutions has become an imperative of certain governments. Moreover, according to North (1990) and Doumbia (2018), the role of institution is to establish certain stable structure of human relations or interactions. Thus, it follows a complex process of essentially changing rules so that they are dynamic over time. In Africa for example, this dynamic is confronted with traditions and codes of conduct that have remained more or less and closed in themselves. Fight against poverty, inequality and all forms of social exclusion that this concept implies has since given it legitimacy, and its inclusion in international agenda and national development strategies (Nkamleu, 2017). Thus, inclusive growth is essential for restoring public confidence in the capacity of democratic institutions, technological progress and international economic integration to support greater progress to support and well-being for all (Cordemans, 2019).

According to Siyakiya (2017), the poor quality of institutions has a negative impact on the economies of poor countries and some developed countries in terms of transaction costs while increasing the decision to invest, focusing on areas that are likely to be productive by directing economic activity to productive areas and finally, building up trust and cooperation. In fact, institutional economics explains why developing countries remain poor because of their poor or less efficient institutions. In general, developing countries have weak institutions and fail to support productive investments and protect ownership rights. As a result, some poor countries enrich other countries, contributing to increased inequality (Fosu, 2017). In such cases, society will be able to achieve inclusive growth if it reforms the quality of its institutions to make them strong and to achieve poverty reduction. According to the literature, there is a close link between ownership rights and inclusive growth. This is what prompted Acemoglu et al. (2005) to argue that the
inclusion of new middle class in inclusive economic activity (Ata in economic activity, 2018) finds institutional improvements, human capital, and economic growth. In this regard, Doumbia (2018) finds that only institutional indicators such as effective governance and the rule of law promote inclusive growth that compromised that of other individuals.

It is generally accepted that the quality of institutions is a major determinant of the level of development (Rodrik, 2000; Collier, 2006). Looking at the quality of institutions through the lens of the six indicators of Kaufmann et al. (2005): participation and accountability, political stability, effectiveness of public power, quality of regulation, rule of law, and control of corruption, it is important to note that the weak performance of poor countries in this area, constrains seriously the inclusiveness of their economic growth. In this regard, Welch and Nuru (2006) point out that democratic governance broadens the range of options for human development. Despite recent improvements, least developed countries are facing the major challenge of creating and strengthening the institutions key of competitive democratic governance that can accommodate the objectives of authority and social inclusion (Gerring et al., 2005). Numerous empirical analyses establish the correlation between political stability and economic growth (Alesina and Perotti, 1996).

This relationship appears to be a bidirectional causal effect: on the one hand, inequality increases social discontent which in turn can lead to violent protest movements (Schock, 1996). And on the other hand, political instability, depending on the degree it has reached, can reduce the spread of the effects of growth inclusiveness by disarticulating the state apparatus and social services (health and education in particular), destroying socio-economic infrastructures, weakening territorial integrity, displacing populations, spreading diseases and reducing the agricultural population (FAO, 2005). In political stability, the large number of underdeveloped countries, particularly in sub-Saharan Africa, is experiencing political and military tensions that reduce the inclusiveness of economic growth as a result of their spillover effects. The effectiveness of public power, including the quality of public spending, especially that allocated to social sectors from which the poor are most likely to be excluded. Corruption affects significantly inclusiveness of economic growth in terms of both wealth creation and equitable redistribution (Gyimah-Brempong, 2001; Dincer and Gunap, 2005) and multidimensional social welfare (Gupta et al., 2002; Aidt, 2010). Directly,
corruption deprives state of important resources to support public action in favor of inclusiveness through the financing of education, health and socio-economic infrastructures or simply reduces the effectiveness of programmed that are supposed to benefit the most vulnerable (Olkens, 2005). Corruption also weakens governance and social justice. It also discourages investment (Asiedu and Freeman, 2009) and therefore affects economic dynamism and, even worse; it can at the same time aggravate exclusion of poorest and most vulnerable from the labor market. In terms of corruption, the situation in poor countries is worrying. As mentioned above, there is little work on inclusive growth and its institutional factors, but a few examples can be listed.

Abbe (2019) attempts to explain the role of institutions in the relationship between unemployment and inclusive growth in the ECOWAS zone (Economic Community of West African States), from 2002 to 2016. For this, he uses Kaufman indices as a measure of institutional factors and Ali and Son (2007) method to measure inclusive growth. Applying the Generalized System of Moments Method (system-GMM), he finds that unemployment has a negative but negligible impact on inclusive growth in the Economic Community of West African States. However, the interaction between the quality of all institutions and unemployment is positive on inclusive growth. Studies conducted in the late 1990s on inclusive growth emphasized the importance of economic freedom as a determining factor, particularly in developing countries.

Thus, Kouton (2019) studies the role of institutions in terms of growth inclusiveness in a panel of thirty Sub-Saharan African countries during 1996-2016. To achieve his goal, he used data on economic freedom as an index of institutional factors and GDP per person employed to capture inclusive growth. Using General Moments Methods (GMM) estimation and panel causality testing in a dynamic framework, the results confer a positive and significant effect of economic freedom on inclusive growth. Thus, evidence is provided for the causal relationship between economic freedom and inclusive growth, but not the reverse.

Dounami (2018) uses the Panel Smooth Transition Regression (PSTR) model to examine the importance of good quality institutions (governance) in terms of pro-poor and inclusive growth for a panel of 112 countries. She uses data from World Government Indicators (WGI) proposed by Kaufmann et al. (2005) and the share of poor people to measure inclusive growth. The results show that all indicators of governance only, government effectiveness and the rule of law are favorable to inclusive growth. In sum, all of these authors studied the link between institutions and inclusive growth, but in a multi-country and short-term perspective. They have therefore not taken into account data from the Political Risk Component by International Country Risk Guide and did it in a short period of time. And only Khan et al. (2016) and Abbe (2019) measured inclusive growth and used this index in econometric framework respectively in time series and panel data.

**EMPIRICAL FARMWORK**

This part of this paper presents data and econometric framework. Institutional quality is used to capture the role of democracy and governance and improve the quality of population life. Indeed, a large literature has accumulated to show that macroeconomic stability is not sufficient to lead to economic growth and inclusiveness (Acemoglu, 2008; Keho, 2012). To be effective, the classical factors of economic growth (labor and capital) must be accompanied by residual factors such as the exercise of democracy and political stability.

**Data analysis**

The measure of the effect of institutional factors of inclusive growth has been analyzed using the proxies as four parameters of governance in the Political Risk Component by International Country Risk Guide (ICRG) reports respectively between 1984 and 2018 (Itfikhar and Khalid, 2011). Base on the empirical work, four institutional factors and indicators were selected: (i) Government Stability, (ii) Corruption, (iii) Bureaucratic quality, and (iv) internal conflict. Government stability measures the government's ability to carry out its planned programs and to sustain itself. The corruption indicator measures the extent of corruption and the manner in which public power is exercised for private purposes. The indicator of bureaucratic quality measures the capacity of administration to conduct day-to-day business without major policy changes or disruption to public services. Finally, internal conflict assesses the level of political violence (civil war, coup threat, terrorism, civil unrest) in a country and its actual or potential impact on governance.

These indicators are scored according to the variation in scores and the best institutional quality. They are taken from the International Country Risk Guide database produced by the Political Risk Service Group (PRS Group). Whatever the methodological and statistical reservations about these indicators, they are taken seriously by foreign investors and international organizations. The advantage of these data, unlike other institutional data, is that they extend over a relatively long period from 1984 to 2018 and do not contain missing values. This main advantage gives it a dynamic analysis of the variables and their impact on the economic development and quality of institutions in each country.

In addition, investment is captured by gross fixed capital formation (%GDP), inflation is captured consumer prices (%annual) and life expectancy at birth is the determinants of growth used as control variables in this study. Investment, inflation and life expectancy at birth are taken from the World-Wide Indicators database.

Moreover, exciting results are obtained when the investment equation is introduced into the economic growth equation (Kouton, 2019). Furthermore, according to Hur (2014), in economies where the majority of the population is poor, as is the case in Africa, strategies that promote greater inclusiveness and investment have become more than a necessity. Moreover, the notion of inclusive growth has a favorable echo in the institutions in charge of development and access to socio-economic infrastructure. Sustainable investment policies can directly or indirectly create jobs and ultimately increase economic growth, making it more inclusive. One possible transmission channel for this purpose is the financing of infrastructure projects at the macroeconomic level. These projects can enable states to create additional jobs.
Inflation is the loss of the purchasing power of money that allows for macroeconomic recovery. Indeed, according to the UNDP definition, the redistribution of the benefits of inclusive growth is not obvious and disappears in an environment of macroeconomic instability. In other words, a stable macroeconomic environment is a favorable condition for inclusive growth. It is along the same lines that Kumah and Sandy (2013) found that countries that maintain a stable macroeconomic environment and have a life expectancy that is not too high have initiated inclusive growth thanks to the political and structural reforms that have been implemented (James et al., 2017). Life expectancy at birth shows how long, on average, a newborn can live, if current death rates do not change. In fact, life expectancy at birth has risen steadily in most 1970 OECD (Organization for Economic Co-operation and Development) countries and increasing by over ten years, on average 1970 (James et al., 2017).

Inclusive growth measurement

To be sustainable and effective, economic growth must be inclusive. In other words, it must require an income growth and equity, equality of opportunity and protection of markets and labor transitions. In absence of consensus on measurement, inclusiveness of growth can be seen in Ali and Son (2007), Anand et al. (2015) and ABBE (2019) who applied in an econometric model. Ali and Son (2007) started from a utilitarian social function integrating both dimensions of growth and equity in a unified framework to measure inclusiveness of growth. This function based on a generalized concentration curve, the social mobility curve, such as follows:

\[
S^* = \left\{ y_1, \frac{y_1 + y_2}{2}, \frac{y_1 + y_2 + y_3}{3}, \ldots, \frac{\sum_{i=1}^{n} y_i}{n} \right\}
\]

(1)

Where \( n \) persons in the population with incomes are \( y_1, y_2, \ldots, y_n \), \( y_1 \) is the poorest person and \( y_n \) is the richest. The study used social mobility curve \( S^* \) to calculate an index called Social Mobility Index (SMI), defined as follows: \( \bar{y}^* = \int_{0}^{100} \frac{y}{d} \). The greater \( \bar{y}^* \) is, the greater will be the income. If people have same income (perfect or equitable distribution of income), then \( \bar{y}^* = \bar{y} \). An Equity of Income Index (EI1) defined as follows:

\[
\omega = \frac{\bar{y}^*}{\bar{y}}
\]

(2)

Similarly, the equity of income index equals to 1 when income distribution is perfectly equitable (that is, everyone has same income, \( \bar{y} \)), and 0 when income is inequitable (that is only one person holds the totality of national income) was obtained by re-ordering Equation 2: \( \bar{y}^* = \omega \bar{y} \)

Deriving Equation 2 gives Equation 3

\[
d\bar{y} \frac{\bar{y}}{y} = \alpha d\bar{y} + y d\omega
\]

(3)

Equation 3 shows the change in the social mobility index is a weighted average of the in income and equity trends. In other words, inclusive growth is a weighted average requiring an increase in income and equity where the coefficient of one component matches the level of the other. Thus, when average income (equity) is high, the contribution of the change in equity (income) is higher and vice versa (Ali and Son, 2007).

Inclusive growth \( \bar{y} \), Equation 3 shows the inclusive can be obtained by increasing \( y \), that is increase the income, or by increasing index, or through the combination of both (Abbe, 2019). The reformulation of Equation 3 gives Equation 4 which combines growth and equity into an inclusive growth (as percentage of change of \( \bar{y} \)) as follows:

\[
\frac{d\bar{y}}{y} = \frac{d\bar{y}}{y} + \frac{d\omega}{\omega}
\]

(4)

Therefore, like Anand et al. (2013), Hussein et al. (2017) and Abbe (2019), who apply \( \bar{y} \) the noted growth \( \frac{d\bar{y}^*}{\bar{y}} \) as a proxy for inclusive growth and at the same time as a dependent variable in the regressions. This definition implies that an increase in \( \frac{d\bar{y}^*}{\bar{y}} \) corresponds to greater inclusiveness of growth.

Model specification

Note that the authors who used the measure of inclusive growth in an econometric model have done it on a panel of countries. As mentioned above, empirical work on this subject is non-existent. So, to measure the effect of institutional factors on inclusive growth time series econometric methodology has been employed to avoid the problem of spurious results. In neo-classical production function, the sources of growth are the accumulation of production factors and improvement of aggregate factor productivity. The starting point of this modeling is the Cobb-Douglas production function defined as follow:

\[
Y_{it} = F(A_{it}, L_{it}, K_{it}) = A_{it}K_{it}^{\alpha}L_{it}^{\beta}, \quad \alpha \text{ and } \beta > 0
\]

(5)

Where \( Y_{it} \) is real GDP per capita in the country \( i \) at time \( t \), \( L_{it} \) was a set of labor force; \( K_{it} \) physical capital stock and \( A_{it} \) total factor productivity.

The study reconsider the consumption model in the following autoregressive form:

\[
\frac{d y}{y} = \alpha_0 + \alpha_1 F_t + \alpha_2 X_t + \alpha_3 F_{t-1} + \alpha_4 X_{t-1} + \varepsilon_t
\]

(6)
The ARDL model is the model par excellence for explaining small and long-term effects even for series of different order of integration, as seen with the limit test approach of Pesaran et al. (2001). Finally, the latter have the particularity of integrating temporal dynamics (adjustment delay and expectations) into the explanation of a variable (time series), thus improving forecasts and the effectiveness of policies (decisions and actions), unlike the simple (non-dynamic) model, whose instantaneous explanation (immediate effect or effect not spread over time) only restores part of the variation of the variable to be explained. Another advantage of this approach is that it does not take into account the order of integration of the regressors, unknown or mixed (I(0) or I(1)), which is not possible with conventional cointegration tests (Pesaran and Shin, 1999; Sam et al., 2019). However, it is possible to fall into the degenerate cases of non-cointegration from the ARDL limit test. This possibility has often been ignored in empirical application.

This technique of ARDL bounds testing has been used in several works to solve economic problems in both micro and macroeconomics and more frequently in problems of public spending and taxation (Afonso and Rault, 2009). This Bound test technique provides a perfect explanation of how the variables tested are cointegrated and the order of their integration. Thus, if the case of the lagged independent variable case is met, the ARDL equation is summarized to Dickey-Fuller unit roots and the independent variable is shown as I(0); otherwise, it is I(1). If the test suggests degenerate dependence or non-cointegration, it indicates that the dependent variable is not included in the cointegration equation embedded in this ARDL equation. The motion of the dependent variable does not respond to the motion of the independent variables, again indicating non-cointegration.

RESULTS AND DISCUSSION

The descriptive statistics results

Descriptive statistics can be used to better decide on the reliability of the data. Two important measures are used to check the reliability of the data. One is the measure of trend and the other is the measure of dispersion. Usually, the mean, median and mode are used as a measure of central tendency and the standard deviation, quartile, range and mean deviation are used as a measure of dispersion (Khan et al., 2016). The results show that the mean and median are almost identical, that there is no evidence of skewness, and that almost all variables have small standard deviation, indicating low variation and low consistency in the data (Table 1).

The stationary tests results

To begin, all the variables mentioned in this document were submitted to the tests of Augmented Dickey-Fuller and Philip Perron for their stationarity. Despite the fact that inflation is integrated of order zero, life expectancy, corruption, inclusive growth and internal conflicts are integrated of order 1. In contrast to the last variables,
Table 2. Stationary test of the variables.

| Variable            | ADF       | PP        |
|---------------------|-----------|-----------|
|                     | level     | 1st differenced | level | 1st differenced |
| Inclusive growth    | -3.204 (-4.221) | -3.548 (-5.594) *** | -3.204 (-4.221) | -3.689 (-10.758) *** |
| Bureaucratic quality| -8.067***(-4.252) | -8.067***(-4.252) | -2.808(-4.243) | -8.067***(-4.252) |
| Government Stability| -5.723***(-4.252) | -2.828*(-3.639) | -2.596(-4.243) | -5.716***(-4.252) |
| Internal conflict   | -0.824(-2.632) | -6.652*(-4.252) | -2.382(-4.243) | -6.639***(-4.252) |
| Corruption          | -1.961(-4.243) | -6.007*(-4.262) | -1.976(-4.243) | -8.015***(-4.252) |
| Investment          | -8.067***(-4.252) | -8.067***(-4.252) | -2.808(-4.243) | -8.067***(-4.252) |
| Life expectancy     | -0.292(-4.309) | -6.145*(-4.309) | -0.023(4.252) | -15.601***(-3.646) |
| Inflation           | -6.723***(-4.356) | -4.016*(-4.309) | -4.405*(-4.243) | -6.639(-4.252) |

*, **, *** indicate statistical significance at 10%, 5% and 1% levels. ADF= Augmented Dickey-Fuller; PP= Philip Perron.

Table 3. Result of ARDL bound test.

| H0: No co-integration | Value | 5% critical bounds | 1% critical bounds |
|-----------------------|-------|---------------------|---------------------|
|                       |       | I(0) | I(1) | I(0) | I(1) |
| Computed F statistic  | 4.101 | 1.97 | 3.18 | 2.54 | 3.91 |

Source for critical value: Pesaran et al. (2001). The computed F values at 1% of significance indicating a long run relationship.

Table 4. Long-term estimates (ARDL (1.0.0.1.0.0.0.0) selected based on R-BAR Squared Criterion).

| Regressor               | Coefficient | t-statistic | prob |
|-------------------------|-------------|-------------|------|
| Bureaucratic quality    | -0.359      | -1.084      | 0.288|
| Government stability    | -0.317*     | -2.047      | 0.057|
| Internal conflict       | 0.410       | 1.006       | 0.323|
| corruption              | 0.204       | -0.561      | 0.579|
| investment              | 0.360***    | 6.427       | 0.000|
| Life expectancy         | -0.079      | -1.289      | 0.209|
| inflation               | -0.128*     | -2.911      | 0.007|

The dependent variable is inclusive growth index.

investment, bureaucratic quality and government stability are integrated of order zero and order one respectively. This confirms the hypothesis of Pesaran et al. (2001), which states that the variables should have an order of integration of either I(0) or I(1) to use the Bound test. However, the ARDL Bound test is more preferable to the ARDL because it is applied when the series are stationarily integrated in the same order and are cointegrated, or with an appropriate difference in integration. Also, the Bound test can be used when the series have a mixed order of integration (some being stationary, others non-stationary) but provided that none of the series is beyond I(1) (Pesaran and Shin, 1999; Pesaran et al., 2001). In the end, the limit test is in fact a test of co-integration between integrated series of different orders below I(2). Then, the ARDL test procedures were applied to estimate the long-term relationship (Table 2).

The bounds tests results

Table 3 shows the results of the bounds test. Indeed, the ARDL bounds test is a new technique developed by Pesaran et al. (2001) in order to test the presence of long-term relationships (cointegration) between variables using the Wald test. Since the value of the F-statistic is above the limit, then the null hypothesis of the non-existence of cointegration is rejected. Therefore, there is a long-term relationship between inclusive growth and institutional factors. The next step will be to assess the effects of institutional factors on inclusive growth in Côte d'Ivoire in the long and short term. These results are shown in Tables 4 and 5.
Table 5. Short-term estimates.

| Variable               | Coefficients | t-statistic | prob  |
|------------------------|--------------|-------------|-------|
| Bureaucratic quality   | -0.009       | -0.159      | 0.874 |
| Government stability   | 0.460**      | 2.493       | 0.019 |
| Internal conflict      | 0.244        | 0.638       | 0.529 |
| Corruption             | -0.596       | -1.514      | 0.124 |
| Investment             | 0.340***     | 5.717       | 0.000 |
| Life expectancy        | 0.095        | 0.193       | 0.848 |
| Inflation              | -0.166***    | -4.511      | 0.000 |
| CointE q (-1)          | -1.161***    | -7.644      | 0.000 |

The dependent variable is inclusive growth.

Table 6. Diagnostic tests of ARDL regression result.

| Tests                                | Probability values |
|--------------------------------------|--------------------|
| Breusch-godfrey serial correlation LM test | 0.229***           |
| ARCH Heteroscedasticity tests        | 0.508***           |
| Normality test                       | 2.55***            |

*** denotes acceptance of null hypothesis.

Short and long run contemporaneous estimates

Table 4 shows all the long run estimates in explaining the dependent variable. While government stability is significant at 10% level of significance, the investment is significant at 1% level of significance and the inflation is significant at 5% level of significance. Table 5 shows all the short run estimates in explaining the dependent variable. While government stability is significant at 10% level of significance, the investment and inflation are significant at 1% level of significance. The error cointegration term is negative and significant which means that any exogenous shock in one of the variables will lead to convergence towards the equilibrium. An exogenous shock in the inclusive growth will lead movement towards the original equilibrium every year, thus equilibrium is stable.

The diagnostic test of the ARDL result

Table 6 highlights the results of the tests of error autocorrelation, heteroskedasticity and error normality. The Breusch and Godfrey test allows testing an autocorrelation of order greater than 1 and remains valid in the presence of the lagged endogenous variable among the explanatory variables. Heteroskedasticity qualifies data that do not have a constant variance. Error heteroskedasticity does not prejudice the estimation of the coefficients, but rather the statistical tests since the estimated standard errors of the coefficients are not adequate. The normality test verifies whether the data follow a normal distribution. All three tests show statistically significant results at the 1% threshold, so these values lead us to reject the null hypothesis of the absence of autocorrelation, heteroskedasticity and error normality.

DISCUSSION

The findings from Tables 4 and 5 indicate the long-run and short-term estimate respectively. Unexpectedly, both tables give same results. It reveals that only government stability as institutional factor affects the inclusiveness of growth. So, there is a strong relationship between government stability and inclusive growth. The explanatory variables such as inflation and investment are statically significant and affect the inclusive growth in short and long run. A result in line with the literature shows a positive sign and statistically significant of investment at 1% of significance in affecting positively the growth inclusiveness. Though, this result shows that there is positive relationship between investment and inclusive growth. This necessarily means that investments have a positive influence on the development of the financial sector (Keho, 2012). It indicates that the improvement in investment has resulted in a deepening of the financial sector in the country. For example, coefficient of investment that 1% change in the investment causes 0.360% change in the inclusive growth index in long-term. The result shows a diminishing impact of
investment on inclusive growth. However, this support findings of Kouton (2019), James et al. (2017) and Kumah and Sandy (2013) that said, the impact of investment value added on inclusive growth is still positive. Thus, investment will help to develop the country in order to improve the living conditions of Ivoirians people and better, conducive to inclusive growth. Other result demonstrates that macroeconomic stability is represented by inflation. There is a negative and significant effect of inflation on inclusiveness at 1 and 5% especially in short and long-term. This result suggests that a policy aimed at reducing the rate of inflation will favor the inclusiveness of the country's growth. Thus, higher inflation is associated with less poverty reduction, through lower average welfare growth as well as with an adverse contribution to distributional effects confirm by Kouton (2019). In particular, poor households are usually more affected by food price inflation as they need to spend disproportionately more on food, and substitution possibilities are limited. Therefore, they are generally more affected by inflation (Khan et al., 2016).

According the interpretation that give ICGR (2019), a score of 0 points equates to very high risk. The unexpected sign is the negative sign and is significant of government stability at 10%. This would imply that a 1% decrease in government stability would result in an increase in the inclusive growth rate of 0.317% points. Indeed, this would not mean that government stability negatively affects the inclusiveness of growth, but the political stability process of Côte d'Ivoire has not been accompanied by an improvement of people welfare.

The CUSUM and CUSUMQ tests dispense with prior knowledge of the date of rupture (Figures 1 and 2). These tests are based on recursive residues. CUSUM
uses cumulative sum of the recursive residuals while CUSUMSQ uses square of the recursive residuals. If the curves exit the corridor stylized by the dotted lines, it is concluded that there is instability in the model. Otherwise, the model can be considered stable over the entire period. Here, none of the CUSUM and CUSUMQ statistics crossed the lines. The study can therefore conclude that the model has remained stable over the entire study period.

Conclusion

An investigation into the measure of institutional factors using International Country Risk Guide on inclusive growth in Côte d’Ivoire is the focus of this study. The study employed a data span of 34 years using the Autoregressive Distributed Lage Model bounds tests and Ali and Son (2007b) method to measure inclusive growth. The study revealed that only government stability as institutional factor affects inclusive growth significantly and negatively in short run but affects positively in long run. Though, this result shows that there is a relationship between government stability and inclusive growth. Inflation and investment affect significantly inclusive growth in short and long run. In view of the above, this study appeals to competent authorities that they should first reform their institutional system before fully embarking on the path of inclusive growth. Clearly, government would gain to set up a government unit, a strong legislation and assistance in the effective population for more inclusive economic growth. The establishment of such institutions is a long and even delicate process.

CONFLICT OF INTERESTS

The author has not declared any conflict of interests.

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