Sex Differences in Responses to Antidepressant Augmentations in Treatment-Resistant Depression

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Abstract

Background: Women are nearly twice as likely as men to suffer from major depressive disorder. Yet, there is a dearth of studies comparing the clinical outcomes of women and men with treatment-resistant depression (TRD) treated with similar augmentation strategies. We aimed to evaluate the effects of the augmentation strategies in women and men at the McGill University Health Center.

Methods: We reviewed health records of 76 patients (42 women, 34 men) with TRD, treated with augmentation strategies including antidepressants (AD) with mood stabilizers (AD+MS), antipsychotics (AD+AP), or in combination (AD+AP+MS). Clinical outcomes were determined by comparing changes on the 17-item Hamilton Depression Rating Scale (HAMD-17), Montgomery-Åsberg Depression Rating Scale (MADRS), Quick Inventory of Depressive Symptomatology (QIDS-C16), and Clinical Global Impression rating scale (CGI-S) at the beginning and after 3 months of an unchanged treatment. Changes in individual items of the HAMD-17 were also compared between the groups.

Results: Women and men improved from beginning to 3 months on all scales (P<.001, ηp² ≥ 0.68). There was also a significant sex × time interaction for all scales (P<.05, ηp² ≥ 0.06), reflecting a greater improvement in women compared with men. Specifically, women exhibited greater improvement in early (P=.03, ηp²=0.08) and middle-of-the-night insomnia (P=.01, ηp²=0.09) as well as psychomotor retardation (P<.001 ηp²=0.16) and psychic (P=.02, ηp²=0.07) and somatic anxiety (P=.01, ηp²=0.10).

Conclusions: The combination of AD+AP/MS generates a significantly greater clinical response in women compared with men with TRD, supporting the existence of distinct pharmacological profiles between sexes in our sample. Moreover, they emphasize the benefit of augmentation strategies in women, underscoring the benefit of addressing symptoms such as insomnia and anxiety with AP and MS.

Keywords: Antidepressants, antipsychotics, mood stabilizers, major depressive disorder, treatment-resistant depression
Significance Statement

Women are nearly twice as likely as men to suffer from major depressive disorder. Yet, there is a dearth of studies comparing the clinical outcomes of women and men with treatment-resistant depression treated with similar medication. We compared the improvement of women and men treated with similar combinations of medication (antidepressants with mood stabilizers and/or antipsychotics) at the McGill University Health Center. We found that the depressive symptoms of women and men improved significantly over 3 months. We also found that women improved more than men this period. Specifically, the use of mood stabilizers and/or antipsychotics in women improved insomnia, anxiety, and psychomotor retardation more than in men. Our results support the existence of distinct pharmacological profiles between sexes. Moreover, they emphasize the benefit of augmentation strategies in women, underscoring the benefit of addressing symptoms such as insomnia and anxiety with antipsychotics and mood stabilizers.

Methods

The present naturalistic study conducted at the specialized mood disorder clinic of McGill University primarily aimed to evaluate the use of pharmacological combinations of AD+AP, AD+MS, and AD+AP+MS in male compared with female TRD patients. The secondary objective is to investigate possible differences in sociodemographic, clinical, and treatment patterns between male and female TRD patients.

Introduction

Women are nearly twice as likely as men to suffer from major depressive disorder (MDD), and this sex difference is among the most robust of findings in psychopathology research (Weissman et al. 1996; Kessler et al. 2003; Wilhelm et al. 2008; Parker and Brotchie 2010; Salk et al. 2017). Despite the greater prevalence of depression among women, only a few studies have investigated the issue of sex differences and psychopharmacological response in MDD, particularly treatment-resistant depression (TRD) (LeGates et al. 2019; Rubinow and Schmidt 2019; Bartova et al. 2021).

Antidepressants (AD) are the first-line treatment for MDD (NICE 2010; Bauer et al. 2015; Cleare et al. 2015; Kennedy et al. 2016), yet more than 30% of patients show an inadequate response to initial pharmacological treatments (Rush et al. 2006; Berlim and Turecki 2007). International guidelines and clinical studies suggest that MDD non-responding to 2 adequate trials with AD, also called TRD, should be treated with a combination of different classes of AD, or augmentation strategies with antipsychotics (AP) and/or lithium and valproic acid (mood stabilizers [MS]) as well as other treatment modalities (including brain stimulation techniques) (Lam et al. 2009; Ghabrash et al. 2016; Kennedy et al. 2016; Gobbi et al. 2018).

Early evidence showed sex differences in the clinical outcome of augmentation strategies in MDD. For instance, T3 (L-triiodothyronine) was observed to be more effective in the augmentation of AD treatment in women than in men (Altschuler et al. 2001). Additional work in current and novel augmentation strategies may be useful in identifying personalized approaches to optimize treatment in both women and men (LeGates et al. 2019). To the best of our knowledge, clinical response rates to AD and a combination of augmentation strategies with either AP or MS has not been explored comprehensively between women and men. The present naturalistic study conducted at the specialized mood disorder clinic of McGill University primarily aimed to evaluate the use of pharmacological combinations of AD+AP, AD+MS, and AD+AP+MS in male compared with female TRD patients. The secondary objective is to investigate possible differences in sociodemographic, clinical, and treatment patterns between male and female TRD patients.

Patients

Charts of 206 patients meeting DSM-IV criteria for a major depressive episode for ≥2 months were reviewed (American Psychiatric Association, 2000). A total 76 patients met the criteria for TRD by failing ≥2 pharmacological trials with different AD in mono or combination therapy at an adequate dose and for ≥3 weeks (Lam et al. 2009). All patients had at least a mild to severe major depressive episode, suggested by a score of ≥13 on the Hamilton-Rating Scale for Depression (HAMD-17) and a score of ≥20 on the Montgomery–Åsberg Depression Rating Scale (MADRS) based on cut-off values proposed by Zimmerman et al. (2013). Patients were treated with augmentation strategies, including ADs with MS (AD+MS), AP (AD+AP), or both (AD+AP+MS).

Clinical Evaluation

Chart analysis was performed by 2 authors (N.A.N. and G.G.) and evaluated at baseline, before the beginning (T0), and after at least 3 months of an unchanged pharmacological treatment (T3). At T0 and T3, patients were assessed on the following behavioral scales: HAMD-17 (Hamilton 1986), MADRS (Montgomery and Åsberg 1979), the Quick Inventory of Depressive Symptomatology (Rush et al. 2003) (QIDS-C16), and the Clinical Global Impression–Severity of Illness (Rush et al. 2003) (CGI-S). The response was defined as a ≥50% reduction from the pre-treatment in the HAMD-17 score. Remission was defined as a score <7 of the HAMD-17 at T3.

Reliability and Inter-Rater Agreement for Psychometric Scales

The internal consistency was previously assessed utilizing Cronbach’s alpha, and an acceptable reliability was found for all scales (HAMD-17: α = 0.82; QIDS-C16: α = 0.77) (Nuñez et al. 2018). Inter-rater reliability was previously assessed using...
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Cohen’s kappa (Cohen 1968) on a sample of 140 patients (Nuñez et al. 2018) with a moderate to good agreement for all behavioral scales (HAMD-17: κ=0.58; QIDS-C16: κ=0.61; CGI-S: κ=0.72).

**Statistical Analyses**

Group comparisons on patients’ demographics were computed through the Pearson’s chi-square ($\chi^2$) test or by Fisher’s exact test (if n ≤5 in each subgroup). Changes in scales were analyzed using repeated-measures ANOVA with sex as between-subject factor and time as a within-subject factor, followed by Tukey post-hoc analyses. Effect sizes are reported for t tests (Cohen’s d) and ANOVA (partial eta-squared, $\eta^2_p$). Small, medium, and large effect sizes were respectively 0.2, 0.5, and 0.8 for “d”, and 0.01, 0.06, and 0.14 for $\eta^2_p$ (Cohen 1968). Analyses were performed using R Statistical Software (R Core Team, 2020). Significance was set at $P<.05$. Data are presented as mean±SD, except when otherwise specified.

**RESULTS**

**Demographics**

A total 76 patients were included in the study (age: 47.71 ±12.50 years; Table 1). Women and men showed a moderate level of resistance based on the Maudsley Staging Method (women: 9.92 ±1.89, men: 9.47 ±1.67). Women previously had tried an average 5.2 (±3.2) medications and men an average 4.4 (±2.0) medications. Pharmacotherapies are described in Tables 2 and 3. At T0, women and men had a moderate/severe depression (HAMD-17: 24.98 ±5.91 and 22.47 ±5.99, respectively). Further characteristics of this sample can be found in our previous work (Nuñez et al. 2018; Moderie et al. 2022).

**Response and Remission**

Response and remission rates of women and men did not differ significantly between the 2 groups ($P \geq .12$; Table 4). Of note, no suicide attempt or suicidal behavior occurred during the 3-month follow-up of the patients.

**Clinical Outcomes in Women vs Men**

For the HAMD-17, 2-way repeated-measures ANOVA indicated a significant sex × time interaction ($F_{1,65}=8.55, P=.005, \eta^2_p=0.11$; Figure 1) as well as significant main effect of time ($F_{2,130}=167.5, P<.001, \eta^2_p=0.71$). There was no main effect of sex ($F_{1,65}=0.21, P=.64, \eta^2_p=0.01$). For the MADRS, 2-way repeated-measures ANOVA indicated a significant sex × time interaction ($F_{1,65}=3.93, P=.05, \eta^2_p=0.06$) as well as significant main effect of time ($F_{2,130}=144.97, P<.001, \eta^2_p=0.68$). There was no main effect of sex ($F_{1,65}=0.80, P=.37, \eta^2_p=0.01$). For the QIDS-C16, there were significant sex × time interactions ($F_{1,65}=5.40, P=.02, \eta^2_p=0.08$) as well as a significant main effect of time ($F_{2,130}=171.42, P<.001, \eta^2_p=0.73$). There was no main effect of sex ($F_{1,65}=0.65, P=.42, \eta^2_p=0.01$). For the CGI-S, there were significant sex × time interactions ($F_{1,65}=5.47, P=.02, \eta^2_p=0.07$) as well as significant main effect of time ($F_{2,130}=132.48, P<.001, \eta^2_p=0.68$). There was no main effect of sex ($F_{1,65}=0.68, P=.41, \eta^2_p=0.01$). For all scales, Tukey post-host analyses on the sex × time interactions revealed no between-group differences at T0 or T3.

**Table 1. Socio-Demographic and Clinical Characteristics of Patients (Baseline)**

|                     | Women        | Men          | Statistics |
|---------------------|--------------|--------------|------------|
| No. of patients     | 42           | 34           |            |
| Age (y) (mean ± SD) | 47.95 ± 12.24| 47.41 ± 12.99| t=0.18, P=.85, d=0.04 |
| Duration of illness (y) (mean ± SD) | 12.4 ± 12.6 | 16.2 ± 12.3 | t=1.72, P=.09, d=0.38 |
| Place of birth      |              |              |            |
| Africa              | 2 (5%)       | 4 (12%)      |            |
| North America       | 28 (65%)     | 21 (62%)     |            |
| Central or South America | 2 (5%) | 1 (3%) | $\chi^2=20.32, P=.31$ |
| Asia                | 4 (10%)      | 3 (9%)       |            |
| Europe              | 6 (15%)      | 5 (15%)      |            |
| No. of past suicide attempts (mean ± SD) | 0.67 ± 1.28 | 0.18 ± 0.46 | t=2.30, P=.03, d=0.50 |
| No. of past hospitalizations (mean ± SD) | 1.76 ± 1.45 | 1.23 ± 0.93 | t=4.20, P=.18, d=0.50 |
| No. of past medications (mean ± SD) | 5.21 ± 3.19 | 4.41 ± 2.01 | t=1.28, P=.21, d=0.30 |
| MSM (mean ± SD)     | 9.92 ± 1.89  | 9.47 ± 1.67  | t=1.33, P=.19, d=0.25 |
| Depression severity (mean ± SD) |            |              |            |
| HAMD-17             | 24.98 ± 5.91 | 22.47 ± 5.99 | t=1.82, P=.07, d=0.42 |
| MADRS               | 33.19 ± 9.00 | 29.79 ± 7.84 | t=1.75, P=.08, d=0.40 |
| QIDS-C16            | 15.71 ± 3.36 | 14.21 ± 3.81 | t=1.80, P=.07, d=0.42 |
| CGI-S               | 5.38 ± 1.08  | 4.88 ± 1.15  | t=1.93, P=.06, d=0.45 |
| Comorbidities       |              |              |            |
| Patients with anxiety disorders | 26 (62%) | 21 (62%) | $\chi^2=0, P=1.00$ |
| Patients with substance-use disorders | 7 (17%) | 7 (21%) | $\chi^2=0.2, P=.89$ |
| Pharmacological strategy |            |              |            |
| AD+AP               | 20 (48%)     | 15 (44%)     | $\chi^2=0.09, P=.95$ |
| AD+MS               | 9 (21%)      | 8 (24%)      |            |
| AD+AP+MS            | 13 (31%)     | 11 (46%)     |            |
| Psychotherapy       | 21 (50%)     | 11 (46%)     | $\chi^2=1.73, P=.18$ |

aAbbreviations: AD, antidepressants; AP, antipsychotics; CGI-S, Clinical Global Impression rating scale; HAMD-17, 17-item Hamilton Depression Rating Scale; MADRS, Montgomery-Åsberg Depression Rating Scale; MS, mood stabilizers; MSM, Maudsley Staging Method; QIDS-C16, Quick Inventory of Depressive Symptomatology (Clinician-Rated); TRD, treatment-resistant depression.
Table 2. Number of Patients Receiving Different Antidepressants and Respective Doses (mean ± SD)

| Antidepressants | Women n | Mean ± SD | Men n | Mean ± SD |
|-----------------|---------|-----------|-------|-----------|
| Citalopram (mg) | 8       | 33.75 ± 25.04 | 8     | 32.50 ± 11.65 |
| Escitalopram (mg) | 9   | 15.56 ± 5.27 | 3     | 10.00     |
| Fluoxetine (mg) | 2      | 20.00 | 1     | 20.00     |
| Fluvoxetine (mg) | 1    | 50.00 | —     | —         |
| Paroxetine (mg) | 1      | 40.00 | —     | —         |
| Sertraline (mg) | 3      | 141.67 ± 62.92 | 4     | 100.00 ± 40.82 |
| Duloxetine (mg) | 1      | 60.00 | 5     | 54.00 ± 13.42 |
| Venlafaxine (mg) | 12    | 137.50 ± 68.47 | 12    | 146.88 ± 95.22 |
| Desvenlafaxine (mg) | — | — | 1     | 50.00     |
| Bupropion (mg) | 12     | 175.00 ± 58.39 | 7     | 278.57 ± 134.96 |
| Amitriptyline (mg) | —     | — | 3     | 58.33 ± 57.74 |
| Clomipramine (mg) | 1    | 25.00 | 1     | 25.00     |
| Mirtazapine (mg) | 9      | 22.50 ± 13.52 | 5     | 34.50 ± 22.25 |
| Trazodone (mg) | 4      | 81.25 ± 12.50 | 2     | 50.00     |

Table 3. Number of Patients Receiving Different Antipsychotics, Mood Stabilizers and Their Combination and Respective Doses (mean ± SD)

| Antipsychotic | Women n | Mean ± SD | Men n | Mean ± SD |
|---------------|---------|-----------|-------|-----------|
| Aripiprazole (mg) | 3 | 6.7 ± 7.2 | 3 | 1.4 ± 0.8 |
| Olanzapine (mg) | 3 | 9.2 ± 3.8 | — | — |
| Quetiapine (mg) | 11 | 127.5 ± 191.3 | 10 | 107.5 ± 84.2 |
| Risperidone (mg) | 3 | 1.8 ± 1.3 | 1 | 2 |
| Quetiapine (mg)/risperidone (mg) | — | — | 1 | 400/1 |
| Mood stabilizer | — | — | — | — |
| Lamotrigine (mg) | 5 | 47.5 ± 28.5 | 1 | 300 |
| Lithium (mg) | 2 | 450 ± 212.1 | 1 | 600 |
| Topiramate (mg) | 1 | 25 | — | — |
| Valproic acid (mg) | 1 | 125 | 5 | 415.0 ± 213.3 |
| Lamotrigine (mg)/valproic acid (mg) | — | — | 1 | 150/1250 |
| Antipsychotic + mood stabilizer | — | — | — | — |
| Aripiprazole/lithium (mg) | 1 | 5/300 | — | — |
| Aripiprazole/valproic acid (mg) | — | — | 3 | 2/375 ± 216.5 |
| Olanzapine (mg)/lamotrigine (mg)/lithium (mg) | — | — | 1 | 5/50/300 |
| Olanzapine (mg)/quetiapine (mg)/lithium (mg) | — | — | 1 | 5/300/300 |
| Olanzapine (mg)/lamotrigine (mg) | — | — | 1 | 15/600 |
| Olanzapine (mg)/valproic acid (mg) | 2 | 5/250 | — | — |
| Quetiapine (mg)/lamotrigine (mg) | — | — | 1 | 50/600 |
| Quetiapine (mg)/topiramate (mg) | — | — | 1 | 50/25 |
| Quetiapine (mg)/lamotrigine (mg) | 1 | 50/300 | 1 | 600/900 |
| Quetiapine (mg)/lamotrigine (mg) | 2 | 50/122.5 ± 123.7 | — | — |
| Quetiapine (mg)/valproic acid (mg) | 2 | 50/625 ± 530.3 | 2 | 125 ± 35.4/250 |
| Quetiapine (mg)/lamotrigine (mg)/valproic acid (mg) | 1 | 300/300/125 | — | — |
| Risperidone (mg)/valproic acid (mg) | 1 | 125/25/125 | — | — |
| Risperidone (mg)/lamotrigine (mg) | 1 | 0.5/25 | — | — |

Therapeutic range: Aripiprazole [2–15 mg]; Olanzapine [5–20 mg]; Quetiapine [50–300 mg]; Risperidone [0.25–3 mg]; Lamotrigine [25–200 mg]; Lithium [600–1200 mg, based on therapeutic serum levels]; Topiramate [N/A]; Valproic Acid [N/A, based on therapeutic serum levels].

Improvement in Individual Items of the HAMD-17 in Women vs Men

In Table 5, we compared 3-month changes in individual items of the HAMD-17 in women and men. Women exhibited greater improvement in both early ($P = .03, \eta^2 = 0.08$) and middle-of-the-night insomnia ($P = .01, \eta^2 = 0.09$) as well as retardation ($P < .001, \eta^2 = 0.16$), psychic ($P = .02, \eta^2 = 0.07$), and somatic anxiety ($P = .01, \eta^2 = 0.10$). There was a marginal finding for a greater improvement in general somatic symptoms in women vs men ($P = .07, \eta^2 = 0.04$). No significant findings were seen in other clinical scales ($P \geq .15$).

DISCUSSION

This is the first study, to our knowledge, comparing the clinical trajectory of women and men with TRD treated with similar...
One of our main findings is the greater reduction of depressive symptoms in women compared with men (medium effect size) treated with an AD and augmented with an AP and/or MS. Yet, a recent analysis by the Group for the Study of Resistant Depression showed a trend towards a more frequent administration of add-on treatments...
in men than in women (Bartova et al. 2021). Our results emphasize the importance of augmentation strategies in women with TRD. The synergistic effect of AD+AP is well studied in unipolar depression (Dold and Kasper 2017), and preclinical studies have underscored that augmentation with an AP allows for targeting multiple receptors and neurotransmitters systems (Blier and Blondeau 2011). There is also evidence undermining the importance of augmentation strategies including MS for TRD patients (Blier and Blondeau 2011; Dold and Kasper 2017; Gobbi et al. 2018). Historically, women were prescribed more tranquilizing and hypnotic drugs than men, but recently, AP appears to replace the use of those medications (Seifert et al. 2021a,b). Augmentation of AD with evidence-based pharmacotherapies rather than tranquilizing and hypnotic drugs will benefit women with TRD (Kennedy et al. 2016; Seifert et al. 2021b). Sex differences for the prescription of AD and MS among patients with MDD are largely unavailable and remain to be clarified (Seifert et al. 2021b). Our study design did not allow us to systematically address this question, but no significant differences were noted in terms of augmentation strategy. Most patients were augmented with an AP, and quetiapine was the agent most prescribed in both men and women. It is also the medication with the best evidence, as emphasized in a recent Cochrane Review for the management of TRD (Davies et al. 2019). In our study, MS was used less often than AP, without any sex differences. Others have reported a less common administration of MS in women compared with men with MDD (Bartova et al. 2021), a difference likely driven by the contraindication of valproic acid and lithium in women of childbearing age due to their potential teratogenic effects (Gentile 2010; Dold et al. 2016; Munk-Olsen et al. 2018).

The differential treatment outcomes observed between men and women could be explained by a myriad of factors. It has been suggested that women may be more likely to respond to selective serotonin reuptake inhibitors (SSRIs) than a tricyclic AD, whereas men may be more likely to respond to tricyclic AD than an SSRI (Frank et al. 1988; Haykal and Akiskal 1999; Kornstein et al. 2000; Berlanga and Flores-Ramos 2006; Young et al. 2009). The occurrence of certain drug side effects (i.e., weight gain or sexual dysfunction) may also contribute to the differential AD efficacy and tolerability between sexes (Seifert et al. 2021a). Because most patients were treated with SSRIs in our study, this could contribute to explaining the greater improvement in specific symptoms in women compared with men. Nonetheless, there is no definite consensus on whether sex differences in AD efficacy actually exist (Keers and Aitchison 2010; Sramek et al. 2016; LeGates et al. 2019), and the National Institute for Health and Care Excellence explicitly states that little evidence supports prescribing AD according to sex (NICE 2010).

Besides, the pharmacokinetics of augmenting agents might exhibit sex differences with hypothesized differences in drug transporters (Benet et al. 1999), metabolizing enzymes (Harris et al. 1995; Cheung et al. 2006), and resulting plasma levels of medication (Ronfeld et al. 1997; Keers and Aitchison 2010). Pharmacodynamic properties of the augmenting agents may, furthermore, differ in men and women, with distinct effects on neurotransmitter synthesis in men and women (Keers and Aitchison 2010). More studies are needed to elucidate distinct effects of augmenting agents in relation to sex.

The groups in our study were comparable in terms of age, duration of illness, number of past hospitalizations and medications, and comorbidities with substance-use disorders (SUD) and anxiety disorders. Unlike our findings, in patients with MDD (non-TRD), alcohol and drug abuse is more common in men than in women (Marcus et al. 2008). Although the sample size
is limited, our results may indicate that the prevalence of SUD in women with TRD might be higher compared with women with MDD, in line with increased risk for SUD among patients with TRD compared with other depressed patients (Brenner et al. 2019). Notably, most studies in TRD excluded individuals with SUD (Bennabi et al. 2015; De Carlo et al. 2016). Although the European Group for the Study of Resistant Depression did not identify SUD as a risk factor for TRD (Souery et al. 2007), sex differences were not investigated.

Women with MDD (non-TRD) present comorbid anxiety disorders more frequently than men and are more likely to suffer from anxiety prior to the development of depression (Breslau et al. 1995; Yonkers et al. 1996; Howell et al. 2001; Marcus et al. 2005; Grigoriadis and Erlick Robinson 2007; Buxh et al. 2010). Some data even suggest that the increase in anxiety–depression comorbidity may explain the greater lifetime prevalence of depression in women (Breslau et al. 1995). Interestingly, the European Group for the Study of Resistant Depression found no difference in comorbid anxiety disorders when comparing women and men with TRD, which might suggest an attenuation of this sex difference in TRD (Barova et al. 2021). Likewise, we found no differences in comorbid anxiety disorders in women compared with men. However, as shown in the Zurich Cohort Study, women also have higher rates of sub-threshold co-morbid anxiety, which could contribute to the treatment resistance in MDD (Angst and Merikangas 2001; Souery et al. 2007). Our data also align with the findings from the DEPRES I and II studies reporting higher prevalence of insomnia and anxiety symptoms in women compared with men (Angst et al. 2002). Such evidence emphasizes the need to address anxiety and insomnia, particularly in women. AD, particularly SSRIs, may not sufficiently alleviate those symptoms in women (LeGates et al. 2019) and might contribute to the higher number of tranquilizers and hypnotics prescribed for women than for men (Royd et al. 2015; Seifert et al. 2021b). In the current study, augmentation with AP and/or MS helped to significantly reduce (moderate effect size) both insomnia and anxiety in women more than in men.

Another important finding is the larger improvement noted in reported early and middle-of-the-night insomnia in women compared with men. Women with MDD generally report more insomnia symptoms than men (Silverstein 1999; Marcus et al. 2005). Insomnia is an established and modifiable risk factor for depression, the treatment of which offers the critical opportunity to prevent major depressive episodes (Plante 2021). The differential improvement in insomnia in women compared with men was accompanied by large-effect size difference in psychomotor retardation. Although no causality can be drawn, our results suggest that improving sleep with augmenting agents in women could decrease psychomotor retardation.

While we found a distinct clinical improvement on the severity of depression according to the different pharmacotherapy strategies, we did not observe an overall difference in their response or remission rates. Such outcomes should be viewed, considering the long period required to achieve remission or euthymic states in depression (Goodwin et al. 2016). The observed low rates of remission indeed reflect the refractory nature of patients included in this study. However, no suicide or suicide attempts were reported during the study follow-up, underscoring that even if the pharmacological combinations did not lead to remission within 3 months, they may be significant in certain depressive domains such as preventing suicidal behaviors. There was no sex difference observed in the suicidality item of the HAM-D17. As in multiple studies, the number of suicide attempts was higher in women than in men, which could reflect the higher completion rate in men (Kessler et al. 1993; Oquendo et al. 2001). Larger studies are needed to elucidate preferential pharmacotherapy to prevent suicide. The absence of suicide in our cohort can also be linked to follow-up in a tertiary/quaternary clinic with staff fully trained in suicidal prevention and with 24/7 access to psychiatrists and/or psychiatry emergency.

**Limitations**

Several limitations should be considered while interpreting these findings. First, the external validity may be limited by data derived from a university hospital mood-specialized center. Second, we did not match the sample of women and men patients according to single pharmacological agents or dosages as well as to depressive severity. Third, we did not control for the menstrual status/phase of women, which can contribute to the severity of symptoms (Hartlage et al. 2004; Haley et al. 2013; Davari-Tanha et al. 2016; Salk et al. 2017) and AD response. Fourth, the non-blinded retrospective outcome assessments should be considered as well as the limitations of a naturalistic design study. Side effects and adverse events were not systematically documented. Nevertheless, the findings may reflect real-world interactions of clinically selected pharmacotherapies, as clinical treatment was individualized and adjusted to tolerability to favor patients’ preference and positive clinical outcomes (Kennedy et al. 2016; Dold and Kasper 2017). The long follow-up of patients at the clinic also prevents the inclusion of undiagnosed bipolar patients in the sample of TRD (Perlis et al. 2011).

**CONCLUSION**

In our naturalist study in patients with TRD, augmentation strategies generate a significantly greater clinical improvement in women compared with men, supporting the existence of distinct pharmacological profiles between sexes. Moreover, they emphasize the benefit of augmentation strategies in women and highlight the benefit of addressing insomnia and anxiety with AP and MS in this specific population. Further studies linking specific medication and symptoms outcomes in larger sample sizes should provide more insight into these clinical questions to provide personalized management of care of patients suffering from depression. This study paves the way for the investigation of sex differences in TRD, and the data reported here can be used to determine needed sample size in larger trials.

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**Interest Statement**

None.

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