Genetic Interactions with Age, Sex, Body Mass Index, and Hypertension in Relation to Atrial Fibrillation: The AFGen Consortium

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It is unclear whether genetic markers interact with risk factors to influence atrial fibrillation (AF) risk. We performed genome-wide interaction analyses between genetic variants and age, sex, hypertension, and body mass index in the AFGen Consortium. Study-specific results were combined using meta-analysis (88,383 individuals of European descent, including 5,722 with AF). In the discovery analysis, the AF risk associated with the minor rs6817105 allele (at the PITX2 locus) was greater among subjects aged 65 years (interaction p-value = 4.0 × 10−6). The interaction p-value exceeded genome-wide significance in combined discovery and replication analyses (interaction p-value = 1.7 × 10−8). We observed one genome-wide significant interaction with body mass index and several suggestive interactions with age, sex, and body mass index in the discovery analysis. However, none was replicated in the independent sample. Our findings suggest that the pathogenesis of AF may differ according to age in individuals of European descent, but we did not observe evidence of statistically significant genetic interactions with sex, body mass index, or hypertension on AF risk.

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Atrial fibrillation (AF) is a common arrhythmia and is associated with increased risk for stroke, heart failure, and mortality. Previous studies have demonstrated that increasing age, male sex, high blood pressure, and obesity are associated with higher AF risk. AF is heritable, and genetic association studies have identified 16 loci tagged by common genetic variants that are associated with AF.

Typically, genome-wide association studies have assumed that the effect of each tested SNP on AF risk is constant across various risk factors, though some data suggest that the effect sizes may differ for different values of risk factors. For example, variants at the HINTL1 region have been shown to interact with alcohol consumption to affect colorectal cancer risk.
AF across common clinical risk factors could potentially refine our knowledge about the genetic basis of AF in important clinical subsets of individuals. Nevertheless, no large systematic examination of interactions between genetic variants and clinical AF risk factors has been conducted. We therefore aimed to determine whether common genetic variants interact with age, sex, hypertension, and body mass index to modify AF risk in a large sample of individuals of European ancestry.

Results
A total of 88,378 subjects, including 7,292 with AF, were included in the discovery analysis (Table 1). The numbers of included SNPs and values of genomic inflation factors ($\lambda$) for each study (after applying quality control criteria for SNP exclusions) are displayed in Supplemental Table 1. Overall, genomic inflation factors ranged from 0.85 to 1.28–1.47, $p$-value $= 6.2 \times 10^{-10}$ for subjects $\leq 65$ years; $OR = 1.09$, 95% CI 1.03–1.15, $p = 1.4 \times 10^{-3}$ for subjects $>65$. We did not observe any significant interactions between AF-associated SNPs and sex, hypertension, or body mass index.

Interactions with risk factors at known AF loci. We first evaluated the associations between genetic interactions and clinical factors (age, sex, hypertension, and body mass index) with AF at 16 established AF susceptibility loci from prior genome-wide association studies (Supplemental Table 2; significance threshold $= 6.25 \times 10^{-8}$, see methods for explanation). We observed significant interactions with age for SNP rs6817105 (upstream of PITX2 at chromosome locus 4q25; interaction $p$-value $= 4 \times 10^{-3}$; Table 2). The minor C allele of SNP rs6817105 was associated with a greater risk for AF among individuals 65 years of age or younger [odds ratio (OR) = 1.75, 95% CI 1.61–1.91, $p = 6.2 \times 10^{-36}$], than among participants older than 65 years (OR = 1.38, 95% CI 1.28–1.47, $p = 6.3 \times 10^{-17}$). Among other known AF loci, SNP rs3807989 at the CAV1 locus displayed a nominal interaction with age that was not statistically significant (interaction $p = 2.9 \times 10^{-3}$; Table 2). However, the major G allele was associated with higher AF risk in the younger group (OR = 1.25, 95% CI 1.16–1.34, $p = 3.6 \times 10^{-10}$ for subjects $\leq 65$ years; OR = 1.09, 95% CI 1.03–1.15, $p = 1.4 \times 10^{-3}$ for subjects $>65$). We did not observe any significant interactions between AF-associated SNPs and sex, hypertension, or body mass index.
Interactions with risk factors in genome-wide analyses. Table 3 displays the results for SNP interactions with AF risk factors across the genome. The most significant genetic interaction that exceeded our genome-wide significance threshold (an interaction p-value $< 4 \times 10^{-8}$, see methods for explanation) was observed for SNP rs12416673 with body mass index (interaction $p = 2.9 \times 10^{-8}$; 6.4 kb upstream of COL13A1 at chromosome region 10q21; Table 3; Supplemental Figure 2). Specifically, with each 1-unit increase in body mass index, each copy of the minor A allele of SNP rs12416673 was associated with an increased risk for AF (interaction $\beta = 0.0224$, interaction $p = 2.9 \times 10^{-8}$). Additionally, we observed 8 loci that exhibited suggestive interactions with AF risk factors (i.e., the interaction p-value was $< 1 \times 10^{-6}$ for the top SNP, and two or more SNPs in the
Table 2. Multiplicative SNP interactions with AF risk factors at known AF loci. The significance threshold was 0.01/16 = 6.25 × 10⁻⁴. Abbreviations: AF: atrial fibrillation; A1: allele 1; the risk allele was defined based on a prior GWAS; A2: allele 2; A1 freq: allele 1 frequency; Loc: locus; p: P-value for the interaction between the risk factor and the SNP. *Interaction β was from regression using an additive model. Interaction β (se) was calculated as the meta-analysis log(effect) in subjects ≤ 65 years of age minus the meta-analysis log(effect) in subjects >65 years of age, or as the multiplicative interaction between SNP*risk factor for sex (females vs. males), hypertension (hypertensive vs. not), and body mass index (per 1 unit increment).

Table 3. Discovery and replication analysis results of top SNP interactions with AF risk factors. Abbreviations: AF: atrial fibrillation; A1: allele 1; the risk allele was defined based on a prior GWAS; A2: allele 2; A1 freq: allele 1 frequency; Loc: locus; p: P-value for the interaction between the risk factor and the SNP. *Interaction β was from regression using additive model. Interaction β (se) was calculated as the meta-analysis log(effect) in subjects ≤ 65 years of age minus the meta-analysis log(effect) in subjects >65 years of age, or as the multiplicative interaction between SNP*risk factor for sex (females vs. males), hypertension (hypertensive vs. not), and body mass index (per 1 unit increment). ^Known AF loci.
Replication. In total, we selected 10 SNP interactions (Table 3) for replication association testing in four independent cohorts (131,441 individuals, including 5,722 with AF). Only one interaction remained significantly associated with AF. SNP rs6817105 at the 4q25 locus exhibited a significant interaction with age ($p = 9.5 \times 10^{-4}$). As in our discovery analysis, among individuals with the minor C allele of rs6817105, those $\leq 65$ years old had a greater risk for AF (OR = 1.80; 95% CI 1.67–1.95, $p = 6.6 \times 10^{-52}$), than participants older than 65 years (OR = 1.45; 95% CI 1.30–1.61, $p = 1.4 \times 10^{-11}$). Similarly, rs6817105 was associated with a 27% higher AF risk in subjects $\leq 65$ years of age (compared with subjects $> 65$ years of age) in the combined discovery and replication analysis ($p = 1.7 \times 10^{-8}$; Fig. 2). A greater risk of AF for the rs6817105 C allele was observed in participants aged 65 years or younger (OR = 1.78; 95% CI 1.68–1.89, $p = 5.6 \times 10^{-86}$) than in participants older than 65 years (OR = 1.40; 95% CI 1.32–1.49, $p = 7.8 \times 10^{-27}$).

Power calculation. Given the lack of observed associations between SNP interactions with clinical risk factors and AF, we performed power calculations to estimate power for discovery using Quanto24 (http://biostats.usc.edu/Quanto.html; Fig. 3). As an example, we estimated power to observe a SNP interaction with sex, assuming a population comprised of 50% males, an AF population prevalence of 1%, 50% males, SNP marginal effect odds ratios of 1.5, sex marginal effect odds ratios of 1.5, case:control ratios of 1:10, and $\alpha = 4 \times 10^{-4}$. Power calculations were performed using Quanto24.
of interaction was performed. Greater effect sizes of other AF susceptibility SNPs at the 4q25 locus were also observed among younger rather than older individuals in some, but not all, cohorts in a large replication study25. Moreover, in keeping with our observations, prior studies did not find evidence that AF risk is modified by interactions between SNPs at the 4q25 locus and sex26. Our findings demonstrate that genetic variation at the 4q25 locus is, on average, associated with greater risks for early-onset AF.

Our findings are consistent with epidemiologic observations demonstrating greater heritability for earlier onset of AF24. The stronger effect of the 4q25 locus on AF in a younger population implies that the contribution of this locus to AF susceptibility may be more relevant to those with early-onset AF, rather than later onset forms. Overall, our observation that genetic variation at the 4q25 locus is associated with AF (beyond genome-wide significant thresholds) in both younger and older individuals underscores the predominant role of this locus in AF pathogenesis—regardless of age.

PITX2 is a homeobox transcription factor involved in specification of pulmonary vasculature29, cardiac laterality27, and suppression of a left atrial sinoatrial-node like pacemaker28. Heterozygous null Pitx2c mouse hearts are more susceptible to pacing induced AF than are wild-type counterparts25. The relative roles of PITX2 regulation in AF susceptibility in both human development and in adult life are unclear. Future larger studies are warranted to systematically determine whether there are different age-specific etiologic subtypes of AF, and whether PITX2 modulation varies with age according to genotype.

Although our analysis was not designed to specifically quantify the contribution of genetic factors to AF heritability, the absence of observed interactions between AF and sex, body mass index, and hypertension suggests that common variant interactions with these clinical risk factors are unlikely to explain a substantial proportion of variance in AF susceptibility. Larger studies will be necessary to accurately quantify the contributions of both common and rare variation, epigenetic mechanisms, copy-number variation, epistatic effects, and other environmental interactions that may influence AF heritability. Moreover, further examination is needed to determine the extent to which the 4q25 locus, the predominant susceptibility locus for AF, explains the heritability of the condition.

Our study should be interpreted in the context of the study design. First, we included individuals of European ancestry only, so our finding may not be generalizable to other racial groups. Second, AF risk factors were available only at the time of AF onset in case-control studies, rather than before AF onset, potentially biasing toward the null any biologically relevant SNP by risk factor interactions that may occur years before the onset of AF. However, we suspect that such misclassification of risk factor status is unlikely to have resulted in systematic bias for body mass index (which tends to be relatively stable over time30) and age (because an interaction with age and the 4q25 locus is supported by prior observations). Third, our sample size provided limited power to identify interactions with relatively small effect sizes. Additionally, the use of more powerful statistical approaches31, non-multiplicative interactions, and inclusion of additional AF risk factors, may facilitate identification of loci at which genetic interactions exist in relation to AF. Fourth, our single SNP interaction approach does not exclude a lack of interaction with polygenic susceptibility to AF. Fifth, we acknowledge that AF may be clinically unrecognized, leading to misclassification of AF status, and that we lacked power to analyze AF subtypes separately. Future analyses with additional arrhythmia outcomes may help clarify the role of genetic interactions with risk factors across a range of arrhythmia phenotypes.

In summary, we identified a significant interaction with age at the AF susceptibility locus on chromosome 4q25 upstream of PITX2 in individuals of European ancestry. Despite several suggestive SNP interactions with common AF risk factors in discovery analyses, we did not observe substantial evidence for such interactions as common mechanisms underlying AF risk.

Methods

Study population. Discovery cohorts included the: German Competence Network for Atrial Fibrillation and Cooperative Health Research in the Region Augsburg (AFNET/KORA); Age, Gene/Environment Susceptibility Reykjavik Study (AGES) study; Atherosclerosis Risk in Communities (ARIC) study; Vanderbilt electronic medical record-linked DNA repository (BioVU); Cleveland Clinic Lone AF study (CCAF); Cardiovascular Health Study (CHS); Framingham Heart Study (FHS); Ludwigshafen Risk and Cardiovascular Health (LURIC) study; Multi-Ethnic Study of Atherosclerosis (MESA); Massachusetts General Hospital Lone AF study and Myocardial Infarction Genetics Consortium (MGH/MIGEN); Prevention of Renal and Vascular End-stage Disease (PREVEND) study; PROspective Study of Pravastatin in the Elderly at Risk (PROSPER); Rotterdam Study (RS); Study of Health in Pomerania (SHIP); and Women’s Genome Health Study (WGHS). Replication studies included the: Basel Atrial Fibrillation Cohort Study (Beat-AF); Finnish Cardiovascular Study (FINCAVAS); Malmo diet and cancer study (MDCS); and UK Biobank. Detailed descriptions of each study have been previously reported (Supplemental Methods and Supplemental Table 3).

The study protocol was approved by the Ethical Committee/institutional review boards of Ludwig Maximilian University of Munich, National Bioethics Committee, Johns Hopkins Bloomberg School of Public Health, University of Minnesota, Vanderbilt University Medical Center, Cleveland Clinic, University of Washington, Boston University Medical Campus, Rhineland-Palatinate State Chamber of Physicians, Massachusetts General Hospital, University Medical Center Groningen, Leiden University Medical Center, Erasmus MC - University Medical Center Rotterdam, University Medicine Greifswald, Brigham and Women’s Hospital, ethics committee northwest/central Switzerland, ethics committee Zurich, Pirkanmaa Hospital District, and Lund University. All MESA study sites received approval to conduct this research from local institutional review boards at: Columbia University (for the MESA New York Field Center), Johns Hopkins University (for the MESA Baltimore Field Center), Northwestern University (for the MESA Chicago Field Center), University of California, Los Angeles (for the MESA Los Angeles Field Center), University of Minnesota (for the MESA Twin Cities Field Center), Wake Forest University Health Sciences Center (for the MESA Winston-Salem Field Center). Written informed
consent was obtained from all study subjects or their proxies (except BioVU, which is a de-identified EMR biorepository and was “opt-out” prior to December 2014). All experiments were performed in accordance with relevant guidelines and regulations.

**AF ascertainment.** Ascertainment of AF and risk factors in each study has been described previously. Detailed descriptions are provided in Supplemental Table 3. We defined prevalent AF as an event that was diagnosed at or prior to an individual’s DNA collection in cohort studies and on the basis of AF ascertainment in case-control studies. We defined incident AF as an event that was diagnosed after DNA collection among participants free of clinically apparent AF at DNA collection in cohort studies. All AF risk factors except age were ascertained at the time of DNA collection. Age was defined at DNA collection or at the date of recruitment in cohort studies, and at time of AF diagnosis (for AF cases) or at time of DNA collection (for controls) in case-control studies.

**Exposure ascertainment.** Sex was defined on the basis of self-report. Participants were classified as having hypertension if the systolic blood pressure was ≥140 mm Hg or the diastolic blood pressure was ≥90 mm Hg at any clinic visit or exam antecedent to DNA collection, or if the participant was receiving treatment with an antihypertensive medication and had a self-reported history of hypertension or high blood pressure at the time of DNA collection (not applicable in ARIC or FHS; Supplemental Table 3). Body mass index was defined as the weight (kg) divided by the height (m) squared. Blood pressure measurements, medication lists, weights, and heights were ascertained according to study-specific protocols. All participants in the discovery analysis were genotyped on genome-wide SNP array platforms (Supplemental Table 4). Imputed genotypes used in our analysis included approximately 2.5 million genetic variants from the HapMap CEU sample (release 22).

**Statistical analysis.** For each individual study, logistic regression (for prevalent AF; for incident AF in MESA and PREVEND only), generalized estimating equations (in FHS to account for related individuals), or Cox proportional hazard regression (for incident AF in prospective cohorts other than MESA and PREVEND) were performed to examine whether AF was associated with interactions between SNP and AF risk factors. For Cox models, person-time began at study baseline, and individuals were censored at death or loss to follow-up. Robust variance estimates were used when feasible. Details of the regression models are described in Supplemental Table 4. All models were adjusted for age (age at baseline for incident AF and age at AF onset for prevalent AF), sex, site (ARIC and CHS), sub-cohort (FHS), study-specific covariates, and population structure, if applicable. SNPs with low imputation quality (R-square < 0.3) or a minor allele frequency < 0.05 were removed from the analysis.

For interaction analyses involving sex, hypertension, and body mass index, main effect terms for each risk factor, as well as multiplicative interaction terms between each SNP and the respective risk factor, were included in the regression models. Regarding analyses of age, nonlinear associations between SNPs and age could potentially go undetected, due to variable distributions of age across the studies in our analysis. Additionally, some studies had only or mostly early-onset/late-onset AF cases, which limited our ability to perform a regression model with dichotomized age in such samples. Therefore, we assessed SNP interactions with age by comparing meta-analysis estimates of associations between each SNP and AF in individuals ≤65 versus >65 years of age (see below). Studies with <100 AF events in each stratum of age were not included, in order to avoid unstable effect estimates.

Estimators for multiplicative interaction terms were meta-analyzed for sex, hypertension, and body mass index analyses in METAL, using an inverse-variance weighted fixed-effects approach with genomic-control correction. For age, we performed an inverse-variance weighted fixed-effects meta-analysis of the estimators for each SNP separately within each age stratum, with genomic-control correction. Estimators were compared using a Z test, as mentioned above.

SNPs with absolute effect sizes ≥3 or SNPs that were available in only one study were excluded from our final results, to minimize the likelihood of spurious false positive findings. For each of the four genome-wide interaction assessments, we employed an experiment-wide two sided alpha threshold of 0.05, which we adjusted for multiple hypothesis testing. We distributed the alpha differentially across the genome, according to a priori hypotheses about interactions between SNPs and AF risk factors. Specifically, we distributed one-fifth of the alpha to each of the 16 most significantly associated SNPs at genome-wide significant loci identified in prior studies (interaction \( p < 0.01/16 = 6.25 \times 10^{-5} \)). The remaining four-fifths of the alpha were distributed evenly across the genome, for an alpha threshold of 4 \( \times 10^{-8} \) (interaction \( p < 0.04/\sim 1,000,000 \) independent tests). Significantly associated SNPs and SNPs with suggestive associations (i.e., an interaction \( p < 0.005 \) at a recognized AF GWAS locus; or an interaction \( p < 1 \times 10^{-8} \) combined with interaction \( p < 1 \times 10^{-5} \) for two additional SNPs within the same ±50 kb region) in the discovery analysis were carried forward for replication testing. In total, we carried forward 10 SNPs for replication testing (see below), and therefore assumed a replication interaction \( p \) threshold of 0.005 (0.05/10 SNPs). The results of replication studies alone, as well as combined with results from discovery studies, were meta-analyzed as described above.

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Author Contributions

L.-C.W. meta-analyzed the data and prepared all figures and tables. L.-C.W. and S.A.L. drafted the manuscript. L.-C.W., K.L.L., M.M.-N., A.V.S., S.T., P.E.W., J.B., J.C.B., L.-P.L., M.E.K., A.M., H.J.L., M.R., S.T., B.P.K., M.D., D.K., D.I.C., M.F.S., M.W., L.J.L., T.B.H., E.Z.S., A.A., G.P., P.L.T., J.C.D., M.B.S., D.R.V.W., J.D.S., B.M.P., N.S., K.D.T., M.K., K.N., G.E.D., O.M., G.E., J.Y., X.G., I.E.C., P.T.E., B.G., N.V., P.M., I.F., J.H., O.H.F., A.G.U., U.V., A.T., L.M.R., S.K., V.G., D.E.A., D.C., D.M.R., M.K.C., S.R.H., E.J.B., T.L., W.M., J.W.J., B.H.S., S.B.F., C.M.A., and S.A.L. participated in the analysis and/or interpretation of the results. All authors have read and approved the final manuscript. K.L.L., E.J.B., T.L., W.M., J.G.S., J.I.R., P.v.d.H., J.W.J., B.H.S., S.B.F., C.M.A., and S.A.L. supervised the study.

Additional Information

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