The genome-wide supported CACNA1C gene polymorphisms and the risk of schizophrenia: an updated meta-analysis

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Abstract

Background: The CACNA1C gene was defined as a risk gene for schizophrenia in a large genome-wide association study of European ancestry performed by the Psychiatric Genomics Consortium. Previous meta-analyses focused on the association between the CACNA1C gene rs1006737 and schizophrenia. The present study focused on whether there was an ancestral difference in the effect of the CACNA1C gene rs1006737 on schizophrenia. rs2007044 and rs4765905 were analyzed for their effect on the risk of schizophrenia.

Methods: Pooled, subgroup, sensitivity, and publication bias analysis were conducted.

Results: A total of 18 studies met the inclusion criteria, including fourteen rs1006737 studies (15,213 cases, 19,412 controls), three rs2007044 studies (6,007 cases, 6,518 controls), and two rs4765905 studies (2,435 cases, 2,639 controls). An allele model study also related rs2007044 and rs4765905 to schizophrenia. The overall meta-analysis for rs1006737, which included the allele contrast, dominant, recessive, codominance, and complete overdominance models, showed significant differences between rs1006737 and schizophrenia. However, the ancestral-based subgroup analysis for rs1006737 found that the genotypes GG and GG + GA were only protective factors for schizophrenia in Europeans. In contrast, the rs1006737 GA genotype only reduced the risk of schizophrenia in Asians.

Conclusions: Rs1006737, rs2007044, and rs4765905 of the CACNA1C gene were associated with susceptibility to schizophrenia. However, the influence model for rs1006737 on schizophrenia in Asians and Europeans demonstrated both similarities and differences between the two ancestors.

1. Background

Schizophrenia is a chronic, disabling brain disease characterized by delusions, hallucinations, and formal thought disorders in addition to a decline in socio-occupational functioning [1]. Studies with twins [2] and adoptive families [3] have shown that genetic factors are an important cause of schizophrenia. The L-type voltage-gated calcium channels play a unique role in behavioral extinction [4], inhibitory learning, and the maturation of adult cognitive function [5]. The two principal pore-forming subunits of these channels expressed in neurons are the α1C and α1D subtypes [6]. The α1C subtype is encoded by the CACNA1C gene, which is considered a risk factor for schizophrenia based on a large genome-wide association study (GWAS) of European ancestry performed by the Psychiatric Genomics Consortium (PGC) [7]. A growing body of research supports a key role for CACNA1C in schizophrenia in Europeans. Ivorra et al. [8] found that the rs1006737 polymorphism of the CACNA1C gene is strongly associated with schizophrenia and bipolar disorder in a Spanish sample. Wolf et al. [9] suggested that the CACNA1C genotype may explain inter-individual differences in the amygdala volume among patients with schizophrenia in the German sample. The amygdala is not only involved in associative learning but also regulates additional cognitive processes, such as memory and attention [10]. Fatima et al. [11] detected a significant difference in the genotype and allele frequencies for the rs4765905 polymorphism between patients and controls, confirming the hypothesis that the CACNA1C gene was associated with schizophrenia in the Pakistani sample. Rs1006737 and rs4765905 are located in intron 3 of CACNA1C gene. And previous study [12] have shown that disease-related SNPs in the CACNA1C gene (including rs1006737 and rs4765905) were proven to be expression quantitative trait loci (eQTLs), which are located in a region interacting with the promoter of CACNA1C, and may regulate the expression of CACNA1C in the brain.

Based on these findings, we were curious to see if the CACNA1C gene had the same effect on schizophrenia in Asians as it did in the Europeans. The meta-analysis of Zheng et al. [13] and Jiang et al. [14] showed that there was no heterogeneity between the CACNA1C rs1006737 polymorphism in East Asians and Europeans. He et al. [15] also showed that rs1006737 was associated with both schizophrenia and major depressive disorder in the Han Chinese sample. Additional rs1006737 meta-analysis showed an association between this CACNA1C polymorphism and schizophrenia in both the Europeans and Asians when the samples were stratified by ethnicity [16]. However, in a follow-up to the top European GWAS hits, The genotyping performed by Takahashi et al. [17] implicated loci in additional schizophrenia family samples from China and Japan and found no association between 12 polymorphisms (e.g., rs4765905 in the CACNA1C gene) and schizophrenia. Consistent with this finding, Hori et al. [18] found no significant difference in the genotype or allele frequency of the CACNA1C rs1006737 polymorphism between schizophrenia patients and controls in a Japanese sample.

In summary, there is no consensus on whether CACNA1C is associated with schizophrenia or if there are differences in susceptibility to schizophrenia between Asians and Europeans. Therefore, we performed an updated comprehensive meta-analysis on the relationship between CACNA1C gene polymorphisms and schizophrenia, which included case-control studies.

2. Methods
2.1. Literature search strategy

Eligible studies were identified by searching two electronic databases (PubMed and the China National Knowledge Infrastructure [CNKI]). PubMed (2011–present) was used to retrieve English studies only, while the CNKI (2013–present) was used to retrieve Chinese studies only. Using these databases, only completed peer-review studies were likely to be included in our research. The last search update was November 2019. The search terms were rs1006737, rs2007044, rs4765905, CACNA1C, and schizophrenia. The studies used in our meta-analysis were based on the following inclusion criteria: (1) included patients with schizophrenia; (2) contained detailed genotype and allele frequencies; (3) included healthy controls; (4) stated that CACNA1C is a susceptible gene for schizophrenia; (5) were case-control studies. Case-control studies were excluded based on the following criteria: (1) no patients with schizophrenia; (2) lacked genotype frequency data; (3) no control population; (4) abstracts, reviews, and meta-analyses; (5) duplicate sample information; (6) data included in the PGC GWAS from 2014.

2.2 Data extraction

Data extraction, publication bias assessment, heterogeneity detection, sensitivity detection, and statistical analysis were performed as previously described [19]. Briefly, data extraction was carried out independently by two authors in strict accordance with the inclusion and exclusion criteria. Disagreements between the two authors were negotiated until a consensus was reached. The following data were extracted: (1) basic information (e.g., first author’s last name, publication year); (2) sample information (e.g., region, ancestor, source of control, mean age of control group, gender index of the case and control groups, and the number of individuals in the case and control groups); (3) the number of genotypes between cases and controls.

2.3. Assessment of publication bias

Publication bias was assessed using funnel plots in which the x-axis represents the log of the risk ratio, and the y-axis represents the standard error of the log of the OR. By observing the symmetry of the funnel plots, one can judge whether publication bias exists in a study. The degree of publication bias was evaluated using the Egger's test [20]. A P-value > 0.05 indicated publication bias.

2.4. Statistical analysis

Pearson’s chi-square test was used to assess the Hardy–Weinberg equilibrium (HWE) for the controls in each study. The strength of the association between the rs1006737, rs2007044, and rs4765905 polymorphisms and the risk of schizophrenia was measured by odds ratios (ORs) with 95% confidence intervals (CIs). The heterogeneity among the studies was tested by the χ²-based Cochran’s Q-test [21] and I² statistics [22]. For the qualitative Cochran’s Q-test, a P-value > 0.1 indicated a lack of heterogeneity. Therefore, the fixed effects model (Mantel-Haenszel method) was used. In contrast, a P-value < 0.1 indicated the existence of heterogeneity, and the random effects model (M-H heterogeneity method) was used [23]. For quantitative I² statistics, I² was the variation between the studies as a percentage of the total variation. The degree of heterogeneity was divided into low (< 25%), moderate (25−75%), and high (> 75%) groups. The allele contrast model (A vs. a) was used to estimate the effect of the risk allele on the risk of schizophrenia (A = risk allele). Subsequently, multiple pairwise comparisons (e.g., AA vs. aa; AA vs. Aa or Aa vs. aa) were used to determine the most appropriate genetic model (A = risk allele). Subgroup analyses were conducted to evaluate the effects of ancestor on the risk of schizophrenia.

Sensitivity analysis was carried out to assess not only the stability and reliability of the combined results of the meta-analysis but also whether the pooled results were affected by a single study. Meta-regression analysis was performed to assess the impact of different variables (mean age of control group and sex indexes) on the analysis.

All statistical tests were performed using Stata version 12.0 (StataCorp LP, College Station, TX, USA). A P value < 0.05 was considered statistically significant (two-tailed).

3. Results

A total of 67 articles were identified from the PubMed and CNKI database searches. Eighteen studies remained after excluding those that did not meet the inclusion criteria. Specifically, these 18 studies included 14 rs1006737 studies (15,213 cases and 19,412 controls), three rs2007044 studies (6,007 cases and 6,518 controls), and two rs4765905 studies (2,435 cases and 2,639 controls). A PRISMA flowchart describing the selection and screening of the studies used in the meta-analysis is shown in Fig.1. The main characteristics of the included studies are presented in Table 1. The allele frequencies and genotype distributions for all the studies are shown in Table 2.
Table 1. The main characteristics of the studies included in the meta-analysis
| Author          | Year | Country | Ancestor | Source of control group | Mean age of control group | Gender index (case) | Gender index (control) | Case/Control |
|-----------------|------|---------|----------|-------------------------|---------------------------|---------------------|------------------------|--------------|
| Fatima [11]     | 2017 | Pakistani | Caucasian | Population based        | 44                         | 0.33                | 0.71                   | 508/300      |
| Lubeiro [24]    | 2018 | Spain    | Caucasian | Population based        | 29.52                     | 0.72                | 0.98                   | 50/101       |
| Mallas [25]     | 2016 | Mixed    | Mixed     | Population based        | 35.79                     | 0.26                | 0.85                   | 63/124       |
| Porcelli [26]   | 2015 | Korean   | Asian     | Hospital based          | 45.36                     | 0.73                | 1.22                   | 176/326      |
| Ivorra [8]      | 2014 | Spain    | Caucasian | Mixed                   | 43.61                     | 0.79                | 0.75                   | 3063/2847    |
| He [15]         | 2013 | China    | Asian     | Population based        | 30.6                      | 0.53                | 0.86                   | 1235/1235    |
| Guan [27]       | 2013 | China    | Asian     | Population based        | 34.2                      | 0.87                | 0.83                   | 1430/1570    |
| Galaktionova [28]| 2013 | Russia   | Caucasian | Population based        | 36                        | 2.24                | 0.90                   | 188/192      |
| Zheng [13]      | 2013 | China    | Asian     | Population based        | 32.4                      | 1.05                | 1.04                   | 5893/6319    |
| Hori [18]       | 2012 | Japan    | Asian     | Population based        | 46                        | 0.82                | 1.93                   | 552/1132     |
| Zhang [29]      | 2011 | China    | Asian     | Population based        | 22.3                      | 0.49                | 0.60                   | 318/401      |
| Nyegaard [30]   | 2010 | Denmark  | Caucasian | Population based        | -                         | -                   | -                      | 976/1489     |
| Bigos [31]      | 2010 | Caucasian | Population based | 33.09                     | 0.230                     | 1.16            | 282/440     |
| Green [32]      | 2009 | UK       | Caucasian | Population based        | -                         | 0.47                | 1.04                   | 479/2936     |

**Rs2007044**

| Author          | Year | Country | Ancestor | Source of control group | Mean age of control group | Gender index (case) | Gender index (control) | Case/Control |
|-----------------|------|---------|----------|-------------------------|---------------------------|---------------------|------------------------|--------------|
| Bustillo [33]   | 2017 | United States | Caucasian | Population based | 36                         | 0.26                | 0.37                   | 53/129       |
| Zhang [34]      | 2018 | China    | Asian     | Hospital based          | 27.14                     | 0.15                | 0.28                   | 53/129       |

**Rs4765905**

| Author          | Year | Country | Ancestor | Source of control group | Mean age of control group | Gender index (case) | Gender index (control) | Case/Control |
|-----------------|------|---------|----------|-------------------------|---------------------------|---------------------|------------------------|--------------|
| Sudesh [35]     | 2018 | India   | Indian   | Population based        | 38.73                     | 1.01                | 0.483                  | 1005/1069    |

**Notes:** Gender index = female/male; Guan's study included the main characteristics of both rs1006737 and rs4765905; Zheng's study included the main characteristics of both rs1006737 and rs2007044.

Table 2. The distributions of the allele frequency and genotype in the included studies
| Author    | Year | Genotype distribution | PHWE | Allele frequency | P | Cases (%) | Control (%) |
|-----------|------|-----------------------|------|-----------------|---|-----------|-------------|
|           |      | AA       | Aa    | aa    | AA   | Aa    | aa    | A   | a  |
| Rs1006737 |      |          |       |       |      |       |       |     |    |
| Lubeiro   | 2018 | 25    | 23    | 2     | 58   | 38    | 5     | 0.70 | 73(73.0) 27(27.0) 154(76.2) 48(23.8) |
| Fatima    | 2017 | 393   | 84    | 17    | 235  | 54    | 9     | 0.01 | 870(88.1) 118(11.9) 524(88.0) 72(12.0) |
| Ivorra    | 2014 | 1417  | 1271  | 293   | 1420 | 1124  | 240   | 0.41 | 4105(68.9) 1857(31.1) 3964(71.2) 1604(28.8) |
| Galaktionova | 2013 | 78    | 85    | 23    | 80   | 90    | 22    | 0.66 | 241(64.8) 131(35.2) 250(65.1) 134(34.9) |
| Nyegaard  | 2010 | 402   | 444   | 130   | 656  | 675   | 158   | 0.42 | 1248(63.9) 704(36.1) 1987(66.7) 991(33.3) |
| Bigos     | 2010 | 120   | 115   | 47    | 191  | 205   | 44    | 0.31 | 355(62.9) 209(37.1) 587(66.7) 293(33.3) |
| Green     | 2009 | 205   | 208   | 66    | 1367 | 1233  | 336   | 0.02 | 618(64.5) 340(35.5) 3967(67.6) 1095(32.4) |
| Mallas    | 2016 | 23    | 30    | 10    | 56   | 51    | 17    | 0.33 | 76(60.3) 50(39.7) 163(65.7) 85(34.3) |
| Porcelli  | 2015 | 153   | 23    | 0     | 301  | 23    | 2     | 0.11 | 329(93.5) 23(6.5) 625(95.9) 27(4.1) |
| He        | 2013 | 996   | 220   | 14    | 1053 | 166   | 9     | 0.39 | 2212(89.9) 248(10.1) 2272(92.5) 184(7.5) |
| Guan      | 2013 | 1061  | 343   | 26    | 1223 | 327   | 20    | 0.72 | 2465(86.2) 395(13.8) 2773(88.3) 367(11.7) |
| Zheng     | 2013 | 5239  | 635   | 19    | 5706 | 597   | 16    | 0.93 | 11113(94.3) 673(5.7) 12009(95.0) 629(5.0) |
| Hori      | 2012 | 480   | 70    | 2     | 1002 | 127   | 3     | 0.63 | 1030(93.3) 74(6.7) 2131(94.1) 133(5.9) |
| Zhang     | 2011 | 280   | 37    | 1     | 357  | 42    | 2     | 0.53 | 597(93.9) 39(6.1) 756(94.3) 46(5.7) |
| Rs2007044 |      |        |       |       |      |       |       |     |    |
| Zhang     | 2018 | 24    | 25    | 4     | 58   | 57    | 14    | 1.00 | 73(68.9) 33(31.1) 173(67.1) 85(32.9) |
| Bustillo  | 2017 | 26    | 23    | 9     | 35   | 21    | 11    | 0.02 | 75(64.7) 41(35.3) 91(67.9) 43(32.1) |
| Zheng     | 2014 | 2797  | 2540  | 559   | 3166 | 2597  | 559   | 0.42 | 8134(77.5) 3658(22.5) 8929(70.7) 3715(29.3) |
| Rs4765905 |      |        |       |       |      |       |       |     |    |
| Sudesh    | 2018 | 579   | 307   | 51    | 668  | 286   | 38    | 0.29 | 1465(78.2) 409(21.8) 1622(81.8) 362(18.2) |
| Guan      | 2013 | 1307  | 360   | 33    | 1195 | 352   | 24    | 0.74 | 2434(71.6) 426(28.4) 2741(87.2) 399(12.8) |

Notes: $P_{HWE}$, $P$-value of the Hardy–Weinberg equilibrium; A, wild-type allele; a, mutant allele.
3.1. CACNA1C Rs1006737 polymorphism

In the present study, A was defined as the risk allele. The allele contrast (A vs. G), dominant (GA + AA vs. GG), recessive (AA vs. GG + GA), codominance (AA vs. GG and GA vs. GG), and complete overdominance (GG + AA vs. GA) models were used to calculate the pooled ORs. All models except for the codominance model (GA vs. GG) were performed using the fixed effects model (M-H) due to the low heterogeneity. In contrast, the codominance model (GA vs. GG) was performed using the random effects model (M-H) due to its high heterogeneity (I^2 = 99%).

After the overall meta-analysis of rs1006737 was conducted, significant differences were observed between rs1006737 and the risk of schizophrenia using the following models: allele contrast model (A vs. G; OR = 1.151, 95% CI = 1.100−1.204, I^2 = 0.0%, P heterogeneity = 0.867, P = 0.000); dominant model (GA + AA vs. GG; OR = 1.169, 95% CI = 1.107−1.234, I^2 = 0.0%, P heterogeneity = 0.786, P = 0.000); recessive model (AA vs. GG + GA; OR = 1.215, 95% CI = 1.085−1.360, I^2 = 0.0%, P heterogeneity = 0.999, P = 0.001); codominance models (AA vs. GG; OR = 1.296, 95% CI = 1.151−1.459, I^2 = 0.0%, P heterogeneity = 0.993, P = 0.000); codominance model (GA vs. GG; OR = 0.064, 95% CI = 0.024−0.169, I^2 = 99%, P heterogeneity = 0.000, P = 0.000); complete overdominance model (GG + AA vs. GA; OR = 0.897, 95% CI = 0.849−0.948, I^2 = 26.1%, P heterogeneity = 0.173, P = 0.000). The main results are presented in Table 3.

Subsequently, subgroup analysis based on ancestor was performed for rs1006737. For the Caucasians, there were seven studies that included a total of 5,546 patients with schizophrenia and 8,305 controls. Rs1006737 was associated with schizophrenia using all but one genetic model (A vs. G, OR = 1.121, 95% CI = 1.060−1.186, P = 0.000; GA + AA vs. GG, OR = 1.127, 95% CI = 1.047−1.213, P = 0.001; AA vs. GG + GA, OR = 1.203, 95% CI = 1.067−1.357, P = 0.003; AA vs. GG, OR = 1.284, 95% CI = 1.131−1.456, P = 0.000; GA vs. GG, OR = 0.279, 95% CI = 0.132−0.587, P = 0.001). The complete overdominance model (GG + AA vs. GA) resulted in an OR = 0.959, 95% CI = 0.891−1.033, and P = 0.272.

The subgroup analysis also included six studies involving Asians with a total of 9,604 patients with schizophrenia and 10,983 controls. Rs1006737 was associated with schizophrenia using the following models: allele contrast model (A vs. G; OR = 1.206, 95% CI = 1.117−1.303, P = 0.000); dominant model (GA + AA vs. GG; OR = 1.219, 95% CI = 1.123−1.323, P = 0.000); codominance model (GA vs. GG; OR = 0.008, 95% CI = 0.004−0.017, P = 0.000); complete overdominance model (GG + AA vs. GA; OR = 0.827, 95% CI = 0.761−0.899, P = 0.000). There was no association observed using the recessive (AA vs. GG+GA; OR = 1.336, 95% CI = 0.922−1.936, P = 0.125) or codominance model (AA vs. GG; OR = 1.384, 95% CI = 0.955−2.006, P = 0.086) models. The main results are shown in Table 4. Neither the mean age of the control group (slope = 0.995, 95% CI = 0.985−1.005, P = 0.265) nor the sex indexes (case group, slope = 0.943, 95% CI = 0.797−1.115, P = 0.455; control group, slope=1.059, 95% CI = 0.829−1.352, P = 0.616) had any significant impact on the results.

3.2. Rs2007044 and rs4765905 polymorphisms of CACNA1C

Allele G of rs2007044 and allele C of rs4765905 were defined as risk alleles. Because relatively few studies related to rs2007044 and rs4765905 were included in the meta-analysis, only the allele model for these two polymorphisms was analyzed. Significant differences between the patients and controls were observed for both rs2007044 (G vs. A; OR = 1.080, 95% CI = 1.023−1.139, P = 0.006) and rs4765905 (C vs. G; OR = 1.225, 95% CI = 1.100−1.364, P = 0.000). The main results are presented in Table 3.
Table 3. The main results of the overall meta-analysis of CACNA1C polymorphisms.

| Genetic model                        | OR   | 95% CI      | P-value | $\hat{\tau}^2$ (%) | $P_{h}$ | Combination method |
|--------------------------------------|------|-------------|---------|---------------------|--------|-------------------|
| Rs1006737                            |      |             |         |                     |        |                   |
| Allele contrast                      | 1.151 | 1.100-1.204 | 0.000   | 0.0                 | 0.867  | Fixed effects model |
| Dominant                             | 1.169 | 1.107-1.234 | 0.000   | 0.0                 | 0.786  | Fixed effects model |
| Recessive                            | 1.215 | 1.085-1.360 | 0.001   | 0.0                 | 0.999  | Fixed effects model |
| Codominance AA vs. GG                | 1.296 | 1.151-1.459 | 0.000   | 0.0                 | 0.993  | Fixed effects model |
| Codominance GA vs. GG                | 0.064 | 0.024-0.169 | 0.000   | 99.0               | 0.000  | Random effects model |
| Complete overdominance               | 0.897 | 0.849-0.948 | 0.000   | 26.1               | 0.173  | Fixed effects model |
| Rs2007044                            |      |             |         |                     |        |                   |
| Allele contrast                      | 1.080 | 1.023-1.139 | 0.006   | 0.0                 | 0.785  | Fixed effects model |
| Rs4765905                            |      |             |         |                     |        |                   |
| Allele contrast                      | 1.225 | 1.100-1.364 | 0.000   | 0.0                 | 0.719  | Fixed effects model |

**Notes:** $\hat{\tau}^2$ represents the variation in OR attributable to heterogeneity. $P_{h}$ represents the $P$-value of the Q-test for heterogeneity.

**Abbreviations:** CI, confidence interval; OR, odds ratio.

Table 4. Subgroup analysis of the association between rs1006737 and the risk of schizophrenia.
### Summary of pooled ORs

| Ancestor | Summary of pooled ORs | Heterogeneity test |
|----------|------------------------|--------------------|
|          | OR        | 95% CI         | P-value | \( I^2 \) (%) | \( P_h \) |
| **Asian** |           |                |        |                |        |
| Allele contrast | 1.206 | 1.117 -1.303 | **0.000** | 0.0 | 0.583 |
| Dominant   | 1.219 | 1.123 -1.323 | **0.000** | 0.0 | 0.484 |
| Recessive  | 1.336 | 0.922-1.936  | 0.125  | 0.0 | 0.939 |
| Codominance AA vs. GG | 1.384 | 0.955-2.006 | 0.086  | 0.0 | 0.932 |
| Codominance GA vs. GG | 0.008 | 0.004 -0.017 | **0.000** | 83.6 | **0.000** |
| Complete overdominance | 0.827 | 0.761-0.899 | **0.000** | 0.0 | 0.434 |
| **Caucasian** |           |                |        |                |        |
| Allele contrast | 1.121 | 1.060-1.186 | **0.000** | 0.0 | 0.964 |
| Dominant   | 1.127 | 1.047-1.213  | **0.001** | 0.0 | 0.919 |
| Recessive  | 1.203 | 1.067-1.357  | **0.003** | 0.0 | 0.987 |
| Codominance AA vs. GG | 1.284 | 1.131-1.456 | **0.000** | 0.0 | 0.893 |
| Codominance GA vs. GG | 0.279 | 0.132-0.587 | **0.001** | 98.0 | **0.000** |
| Complete overdominance | 0.959 | 0.891-1.033 | 0.272  | 0.0 | 0.457 |

**Notes:** \( I^2 \) represents the variation in OR attributable to heterogeneity. \( P_h \) represents the \( P \)-value of the Q-test for heterogeneity

**Abbreviations:** CI, confidence interval; OR, odds ratio

### 3.3. Sensitivity analysis

Following the sequential exclusion of each study, the combined effect obtained by the new meta-analysis was compared to the total effect. No statistically significant differences were observed, indicating that our analysis results were reliable and stable. The range of OR estimates demonstrated that none of the individual studies "reversed" the observed total effect (Table 5).

Table 5. Results of the sensitivity analysis for the **CACNA1C**rs1006737 polymorphism.
Excluded

| Study     | Sample   | OR      | 95% CI               | P-value | P<sub>h</sub> |
|-----------|----------|---------|----------------------|---------|--------------|
| Lubeiro   | Caucasian| 1.1506206| 1.0998443-1.2037411 | 0.000   | 0.815        |
| Fatima    | Caucasian| 1.1546086| 1.1033095-1.2082929 | 0.000   | 0.878        |
| Mallas    | Mixed    | 1.1497633| 1.0989355-1.202942  | 0.000   | 0.826        |
| Porcelli  | Asian    | 1.1485037| 1.0978361-1.2015097 | 0.000   | 0.903        |
| Ivorra    | Caucasian| 1.166579 | 1.1047648-1.2318518 | 0.000   | 0.866        |
| He        | Asian    | 1.1393697| 1.0879437-1.1932266 | 0.000   | 0.981        |
| Guan      | Asian    | 1.1452636| 1.0925845-1.2004827 | 0.000   | 0.847        |
| Galaktionova | Caucasian| 1.1542471| 1.1029066-1.2079774 | 0.000   | 0.863        |
| Zheng     | Asian    | 1.1498204| 1.094684-1.2077338   | 0.000   | 0.815        |
| Hori      | Asian    | 1.1508508| 1.0996418-1.2044445  | 0.000   | 0.814        |
| Zhang     | Asian    | 1.1517017| 1.1007836-1.204975   | 0.000   | 0.821        |
| Nyegaard  | Caucasian| 1.154137 | 1.0994588-1.2115344  | 0.000   | 0.821        |
| Bigos     | Caucasian| 1.1496404| 1.0980188-1.2036888  | 0.000   | 0.818        |
| Green     | Caucasian| 1.151419 | 1.0981367-1.2072866  | 0.000   | 0.814        |

Notes: P<sub>h</sub> represents the P-value of the Q-test for heterogeneity

Abbreviations: CI, confidence interval; OR, odds ratio

3.4. Publication bias

The symmetry of the funnel plots was used to detect the existence of publication bias (Figures 2−9). The Egger's test quantitatively detected symmetry. Due to the lack of studies on rs4765905, the efficacy of the Egger's test was limited, and the symmetry of the funnel plot could not be detected. No publication bias was detected with the allele contrast model for rs2007044 (G vs. A; t = -0.43, P = 0.743) or rs1006737 (A vs. G; t = 0.86, P = 0.407). Furthermore, no publication bias was detected for rs1006737 with the dominant model (GA+AA vs. GG; t = 0.52, P = 0.613), recessive model (TT vs. GG + GT; t = -0.68, P = 0.507), codominance model (AA vs. GG; t = -0.38, P = 0.713), and complete overdominance model (GG + TT vs. GT; t = -0.31, P = 0.762). However, there was a publication bias for the rs1006737 polymorphism with the codominance model (AA vs. GG; t = -3.88, P = 0.002).

4. Discussion

CACNA1C is associated with bipolar disorder [36], autism spectrum disorder [37], major depression [15], and other central nervous system (CNS) disorders [38]. However, the association between the CACNA1C gene and schizophrenia has not been determined. It is also unclear whether the CACNA1C gene has the same effect on schizophrenia in both Asians and Europeans. Therefore, we conducted a comprehensive meta-analysis on the association between the CACNA1C rs1006737, rs2007044, and rs4765905 polymorphisms and schizophrenia. In the overall analysis, rs1006737 was associated with the risk of schizophrenia in all five genetic models, and rs2007044 and rs4765905 were also related to schizophrenia in the allele model, implying that the CACNA1C gene may influence the risk of schizophrenia. This view is consistent with the results of previous meta-analyses [13, 14, 16, 35, 39, 40].
When we conducted a race-based subgroup analysis of rs1006737, we found that the effects of rs1006737 on schizophrenia in the Asians and Europeans had both similarities and differences. According to the results obtained with the allele (A vs. G) and dominant (GA + AA vs. GG) models, the effect of rs1006737 on the risk of schizophrenia in the Europeans and Asians was consistent (i.e., allele A and genotype GA + AA were protective factors against the development of schizophrenia). However, analysis by the recessive (AA vs. GG + GA) and codominant (AA vs. GG) models showed that the genotype GG + GA was only a risk factor for schizophrenia in the European sample. In contrast, according to the complete overdominance model (GG + GA vs. AA), the GA genotype of rs1006737 only reduced the risk of schizophrenia in the Asian sample. These data suggest that the effect of rs1006737 on schizophrenia is ancestrally diverse.

The current study has two limitations. Due to significant heterogeneity ($I^2 = 99.0\%$) and publication bias (Egger's test $P = 0.002$), the codominant model (GA vs. GG) was not reliable and, therefore, was not a valid gene model for evaluating the rs1006737 polymorphism. In addition, there were few studies on the association between rs2007044 or rs4765905 and schizophrenia. Thus, additional high-quality studies are needed to support our analysis.

This meta-analysis study advanced our understanding of the relationship between CACNA1C polymorphisms and schizophrenia compared to previous literature. First, the current study included more comprehensive studies. A recent meta-analysis of the CACNA1C gene and schizophrenia [39] contained nine studies on the association between rs1006737 and schizophrenia. In comparison, the current study included 14 studies on the association between this polymorphism and schizophrenia, including eight articles [11, 13, 15, 18, 27, 29, 30, 32] shared with [39] along with six additional studies [8, 24-26, 28, 31]. Second, compared to most of the meta-analysis on CACNA1C and schizophrenia, the current study not only included studies on rs1006737 and schizophrenia but also studies on the association between two other CACNA1C polymorphisms (rs2007044 and rs4765905) and schizophrenia. Although the study of Xiao et al. [40] also included these three polymorphisms, it only included samples from Asian samples. Because the current study included samples from both Asians and Europeans, it used a richer source of samples for the analysis. Finally, the current study focused on comparing the impact of rs1006737 on schizophrenia in Asian and European samples. Based on this analysis, the influence model of rs1006737 on schizophrenia in Asian and European samples identified both similarities and differences between the two samples.

### 5. Conclusion

The CACNA1C rs1006737, rs2007044, and rs4765905 gene polymorphisms were associated with the susceptibility to schizophrenia. However, the influence model for rs1006737 on schizophrenia in Asians and Europeans demonstrated both similarities and differences between the two ancestors.

### List Of Abbreviations

GWAS – genome-wide association study;

SNP – single nucleotide polymorphism;

PGC – Psychiatric Genomics Consortium;

OR – odds ratio;

95% CI – 95% confidence interval

### Declarations

**Ethics approval and consent to participate:** Not applicable.

**Consent for publication:** Not applicable.

**Availability of data and materials:** All data generated or analyzed during this study are included in this published article.

**Competing interests:** The authors declare that they have no competing interests.

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**Authors’ contributions:** YPL participated in the study design and drafted the manuscript. XW and XX performed the statistical analysis. BJW and JY contributed to the revision of the final manuscript. All authors have read and approved the manuscript.
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Figure 1

Flow diagram for the search and selection of the included studies. A total of 67 relevant English and Chinese studies were retrieved from the PubMed and CNKI databases. Following removal of the studies that did not meet our inclusion criteria, a total of 19 studies were included in the meta-analysis, including 14 rs1006737 studies, three rs2007044 studies, and two rs4765905 studies.
Figure 2

Forest plot of the allele contrast model (A vs. G) of rs1006737. Significant differences between rs1006737 and schizophrenia risk were observed in the allele contrast model (T vs. G), OR = 1.151, 95% CI = 1.100−1.204, P heterogeneity = 0.867, P = 0.000.
Figure 3

Forest plot of the dominant model (GA+AA vs. GG) of rs1006737. Significant differences between rs1006737 and schizophrenia risk were observed in the dominant model (GA+AA vs. GG), OR = 1.169, 95% CI = 1.107–1.234, P heterogeneity = 0.786, P = 0.000.
Figure 4

Forest plot of the recessive model recessive model (AA vs. GG+GA) of rs1006737. Significant differences between rs1006737 and schizophrenia risk were observed in the recessive model (AA vs. GG+GA), OR = 1.215, 95% CI = 1.085−1.360, P heterogeneity = 0.999, P = 0.001.
Figure 5

Forest plot of the complete codominance models (AA vs. GG) of rs1006737. Significant differences between rs1006737 and schizophrenia risk were observed in the codominance models (AA vs. GG), OR = 1.296, 95% CI = 1.151–1.459, P heterogeneity = 0.993, P = 0.000.
Figure 6

Forest plot of the codominance models (GA vs. GG) of rs1006737. Significant differences between rs1006737 and schizophrenia risk were observed in the codominance models (GA vs. GG), OR = 0.064, 95% CI = 0.024–0.169, P heterogeneity = 0.000, P = 0.000.
Figure 7

Forest plot of the complete overdominance model (GG+AA vs. GA) of rs1006737. Significant differences between rs1006737 and schizophrenia risk were observed in the complete overdominance model (GG+AA vs. GA). OR = 0.897, 95% CI = 0.849–0.948, P heterogeneity = 0.173, P = 0.000.
Figure 8

Forest plot of the allele contrast model (A vs. G) of rs2007044. Significant differences between rs2007044 and schizophrenia risk were observed in the allele contrast model (A vs. G). OR = 1.080, 95% CI = 1.023–1.139, P heterogeneity = 0.785, P = 0.006.
Figure 9

Forest plot of the allele contrast model (C vs. G) of rs4765905. Significant differences between rs4765905 and schizophrenia risk were observed in the allele contrast model (C vs. G). OR = 1.225, 95% CI = 1.100–1.364, P heterogeneity = 0.719, P = 0.000.

Supplementary Files

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