Plant defense against a pathogen drives nonlinear transmission dynamics through both vector preference and acquisition

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Abstract. Host defense against vector-borne plant pathogens is a critical component of integrated disease management. However, theory predicts that traits that confer tolerance or partial resistance can, under certain ecological conditions, enhance the spread of pathogens and spillover to more susceptible populations or cultivars. A key component driving such epidemic risk appears to be variation in host-selection behavior of vectors based on infection status of the host. While recent theory has further emphasized the importance of infection-induced host-selection behavior by insect vectors for plant disease epidemiology, experimental tests on the relationship between vector host-selection preference and transmission are lacking. We test how host plant defense—conferring by the PdR1 gene complex—mediates vector host-selection preference and transmission of the pathogenic bacterium Xylella fastidiosa among grapevine cultivars. We confirmed that PdR1 confers resistance against X. fastidiosa by reducing both pathogen population size and disease severity. We found that vector transmission rates to new hosts exhibited unimodal dynamics over the course of infection when both susceptible and resistant were infected and acted as sources of the pathogen. Transmission from susceptible plants initially increased and then declined as insect vectors avoided severely diseased plants. While transmission from PdR1-resistant plants also initially increased and then declined as well, this was not due to avoidance by vectors, although the exact mechanism remains unclear. We show that (1) vector preference changes over the course of disease progression, (2) vector preference is clearly important but a poor predictor of transmission, and (3) the post-latent incubation period—in which plant hosts are infectious but asymptomatic—is likely a key period for vector transmission of X. fastidiosa. Our results suggest that, consistent with theory, defensive traits lengthen the duration of the incubation period, increasing X. fastidiosa transmission. However, defensive traits may over the long-term ultimately reduce spread possibly through induced resistance. Vector host-selection preference, host resistance, and transmission are clearly dynamic, changing over the course of disease progression. Understanding these dynamics is critical for broader insights into the epidemiology of vector-borne plant pathogens, theory development, and deploying disease-resistant cultivars in an effective and sustainable manner.

Key words: condition-dependent movement; Graphocephala atropunctata; host manipulation hypothesis; host resistance; Vitis arizonica; Vitis vinifera.

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INTRODUCTION

Recent emergence events of vector-borne plant diseases have caused widespread socio-ecological harm. Fungal pathogens spread by wood-boring beetles have devastated forests from Hawaiian Ohi’a lehua to North American conifer forests, while viruses and bacteria spread by Hemipteran vectors—such as huanglongbing disease of citrus—exacerbate vulnerability and threaten biodiversity of agricultural systems worldwide (Anderson et al. 2004, Cheatham et al. 2009, Hulcr and Dunn 2011, Savary et al. 2012).

Host defense against pathogens in agricultural crops is one of the most successful and durable strategies to manage agricultural diseases (Gilligan 2008). The particular form of host defense—whether resistance or tolerance—can also influence pathogen spread broadly. Resistance alleviates the fitness costs of infection by reducing the pathogen population in host tissues, whereas tolerance alleviates the fitness costs by ameliorating disease symptoms with little effect on pathogen population size (Cooper and Jones 1983). Variation among hosts in tolerance and resistance can be maintained over evolutionary timescales through ecological feedbacks (Best et al. 2008, Brown and Tellier 2011) and has important epidemiological consequences (Dwyer et al. 1997, Laine et al. 2011, Borer et al. 2016). Generally, tolerance traits are predicted to increase the prevalence of a pathogen within a host population, though the mechanism depends on the precise form of tolerance (Roy and Kirchner 2000, Miller et al. 2006, Best et al. 2008). While variation in host defense may play an important role in vector-borne pathogen spread, this question has received relatively little attention in the literature compared with directly transmitted pathogens. Nonetheless, the relationship between defense and transmission for vector-borne pathogens may be substantially more nuanced because of the complexity of host-selection behavior by vectors (Zeilinger and Daugherty 2014).

Many insect vectors of plant pathogens—including viral, bacterial, and fungal pathogens—exhibit preference for feeding on hosts based on the infection status of the hosts and of the vectors themselves (reviewed in Eigenbrode et al. 2018). A recent flourishing of theory has highlighted how different components of host-selection behavior—namely attraction/orientation vs. leaving/departure behaviors—change based on leaving status and, in turn, influence pathogen spread (Sisterson 2008, Roosien et al. 2013, Shaw et al. 2017, Donnelly et al. 2019). However, while infection-induced host-selection preference by vectors appears to be common and is predicted to be epidemiologically important, few empirical studies have attempted to test the influence of preference on vector transmission of plant pathogens.

A central challenge in testing the relationship between vector preference and transmission is inducing or measuring variation in vector host choice and transmission concurrently. Some studies have shown that host selection by vectors varies dynamically as disease progresses (Blua and Perring 1992, Werner et al. 2009, Legarrea et al. 2015, Lu et al. 2016, Daugherty et al. 2017), varies positively with host pathogen burden (Legarrea et al. 2015, Lu et al. 2016), or coincides with vector acquisition of the pathogen (Legarrea et al. 2015). However, the only study to our knowledge that showed an empirical relationship between vector preference and transmission to uninfected hosts was conducted by Daugherty et al. (2017), who showed that increasing vector avoidance of Xylella fastidiosa-infected grapevines as disease worsened coincided with reduced transmission to healthy host plants across different temperature regimes.

Comparing vector preference and transmission among plant species or genotypes that vary in defense could also offer an opportunity to test for preference–transmission relationships (Legarrea et al. 2015). Tolerant hosts decouple pathogen burden from the phenotypic responses to infection—which are often used by vectors in host selection (Eigenbrode et al. 2018)—and therefore allow for a decoupling of the influences of preference and pathogen population size on transmission. Furthermore, as hypothesized by Zeilinger and Daugherty (2014), the interactions between different forms of defense and vector preference can have important epidemiological implications. Traits conferring resistance against a pathogen should broadly reduce pathogen spread regardless of the particularities of vector preference. On the other hand, pathogen spread can be greatly enhanced when hosts are tolerant.
to infection and when vectors avoid diseased hosts. Understanding the precise interplay between host defense and vector transmission can improve disease management because of their influence on the risk of disease spillover from partially resistant or tolerant hosts into nearby susceptible host populations (Sisterson and Stenger 2018).

*Xylella fastidiosa* (family Xanthomonadaceae) is a xylem-limited bacterium associated with over 500 plant taxa and causes diseases in many crop species, including olive quick decline syndrome, citrus-variegated chlorosis, Pierce’s disease of grapevines, and leaf scorch of almond, oleander, and coffee (European Food Safety Authority [EFSA] 2018). The bacterium is transmitted in a propagative persistent but non-circulative manner by xylem-sap-feeding insect vectors in the family Cicadellidae and superfamily Cercopoidea, with all xylem-sap-feeding Hemiptera regarded as potential vectors (Sicard et al. 2018). Cicadellid vector species consistently avoid *X. fastidiosa*-infected symptomatic plant hosts but show no discrimination between infected asymptomatic hosts and uninfected hosts (Marucci et al. 2005, Daugherty et al. 2011, De Miranda et al. 2013).

Pierce’s disease costs California viticulture ~$100 million per year in yield losses and insecticide use to manage vectors of *X. fastidiosa* (Tumber et al. 2014). Such insecticide use potentially harms biodiversity and the provisioning of agroecosystem services (Cheatham et al. 2009, Chagnon et al. 2015). Alternatively, grapevine cultivars have been developed from hybridizing domesticated grape *Vitis vinifera* with a North American native congener *V. arizonica* (Krivaneck et al. 2006). Putative resistance is conferred by the *PdR1* quantitative trait locus, and while the precise mechanism of resistance is still unresolved, backcrossed lines harbor much lower populations of *X. fastidiosa* and exhibit negligible disease symptoms (Krivaneck and Walker 2005).

While much reduced, population sizes of *X. fastidiosa* in *PdR1* vines can still exceed that required for vector transmission (Hill and Purcell 1997, Krivaneck and Walker 2005). Given that symptom development is slow (Krivaneck et al. 2005) and insect vectors prefer asymptomatic or healthy hosts, *PdR1* grapevines could act as reservoir hosts of *X. fastidiosa*, enhancing vector transmission among vines and spillover to more susceptible vineyards. We hypothesize that transmission from susceptible grapevines will exhibit a unimodal curve—initially increasing as pathogen burden in infected plants increases, then decreasing as disease symptoms become increasingly severe—leading to vector avoidance and reduced acquisition (Fig. 1). At this later stage, vector transmission from tolerant or partially resistant hosts could exceed that from susceptible hosts.

To test our hypothesis on the epidemic risk from *PdR1* grapevines, we experimentally assessed host-selection preference and transmission dynamics by the efficient cicadellid vector *Graphocephala atropunctata* from inoculated *PdR1*-resistant grapevines and closely related susceptible vines. We compared transmission biology between *PdR1* resistant and susceptible vines over multiple time points and employed structural equation modeling to test which components of the vector transmission process best explained variation in transmission. From this, we show that host susceptibility mediates transmission most clearly through differences in pathogen population size and that host-selection by vector appears to play an important but more minor role.

![Fig. 1. Qualitative theoretical predictions on vector transmission dynamics in the context of plant host defense (or lack thereof) and vector avoidance of symptomatic hosts. Here, we assume resistance is partial, meaning that the pathogen can colonize the host but population growth is limited. The decline in transmission from susceptible hosts is predicted to correspond to increasing vector avoidance of symptomatic plants.](www.esajournals.org)
**Materials and Methods**

**Plants and insects**

All resistant and susceptible grapevine genotypes were segregates of a cross between *V. vinifera* cv. Airen and a hybrid of *V. rupestris* × *V. arizonica* (b40-14 background), with resistance conferred by the *PdR1* gene (Krivaneck *et al.* 2006, Walker and Tenscher 2016). We repeated the transmission experiment over two years: In 2016, we used one susceptible genotype (or line) and one resistant genotype; in 2017, we used two susceptible and two resistant genotypes, with one of the susceptible genotypes being the same in both years. See Appendix S1 for more information on experimental methods.

In both years, transmission trials were conducted with green cuttings grown at the Oxford Tract greenhouses at University of California Berkeley according to the methods described in Appendix S1. Colonies of *G. atropunctata* (Signoret; blue-green sharpshooter, Hemiptera: Cicadellidae) were started from wild-collected individuals and all insects were pre-screened for *X. fastidiosa* infection prior to transmission experiments.

To establish *X. fastidiosa*-infected source plants, we mechanically inoculated 3-month-old resistant and susceptible vines near the base of the main stem with 10 μL of a turbid suspension (OD₆₀₀ > 1) of *X. fastidiosa* subsp. *fastidiosa* culture (STL isolate, American Type Culture Collection 700963) in succinate–citrate–phosphate (SCP) buffer as described in Hill and Purcell (1995). Test plants were mock-inoculated with 10 μL of SCP buffer only.

**Transmission experiment**

In both years, we paired one *X. fastidiosa*-free test plant of a susceptible genotype and one inoculated source plant—either *PdR1* resistant or susceptible—in a tent-style BugDorm cage (61 cm³, BioQuip Products, Rancho Dominguez, California, USA). At the beginning of the trials, eight *G. atropunctata* adults were introduced into the middle of the cage, equidistant from the two plants. In 2016, we repeated the trials at 3, 8, and 12 weeks post-inoculation; in 2017, we repeated the trials at 2, 5, 8, and 14 weeks post-inoculation. New plants were used in each trial, making each trial independent. Each combination of genotype and weeks post-inoculation was replicated eight times in each year.

Once the *G. atropunctata* vectors were introduced into the cage, we recorded the number of insects on each plant an in neutral space (cage walls and pots) repeatedly at pre-determined times from the start of the trials. In 2016, these times were as follows: 1 min, 5 min, 10 min, 15 min, 30 min, 45 min, 1 h, 2 h, 4 h, 6 h, 24 h, 30 h, 48 h, 3 d, 4 d, 5 d, 6 d, 7 d, and 8 d. In 2017, we used the same observation times but shortened the total duration of the trials to 4 d, because the 2016 data showed that the *G. atropunctata* reached equilibrium in their movements by then (Appendix S2: Figs. S1, S2).

At the end of the trials, we removed all insects and noted the severity of Pierce’s disease symptoms on the source plants using the 0–5 symptom severity scoring index developed by Guilhabert and Kirkpatrick (2005). We then estimated population sizes of live *X. fastidiosa* in the source plants by serial dilution culturing (Hill and Purcell 1995). To estimate vector acquisition of *X. fastidiosa*, we assayed the infection status of all *G. atropunctata* in trials using qPCR (Appendix S1). To estimate transmission to uninfected test plants, we assayed the infection status of all test plants by culturing 12 weeks after the end of the trials.

**Statistical analysis**

We first tested for differences in Pierce’s disease symptom severity in source plants with a partial odds ordinal logistic regression model, because the response variable—PD symptom severity index—represented an ordinal categorical variable. We analyzed variation in population size of *X. fastidiosa* in inoculated source plants using a generalized linear model with quasi-Poisson distributed error to correct for over-dispersion. While the quasi-Poisson model showed modest nonconstant error variance, it performed much better than Poisson or negative binomial regression models (results not shown). Both tests included genotypes, weeks post-inoculation, and their interaction as explanatory variables.

We tested for preference of *G. atropunctata* vectors for *X. fastidiosa*-infected source plants or uninfected test plants by estimating attraction and leaving rates from each plant. We fit our repeated-measures count data on *G. atropunctata*
location to the consumer movement model developed by Zeilinger et al. (2014) and generalized to multiple consumers per preference trial by Gray et al. (2020) for each combination of plant genotype and weeks post-inoculation. We fit four variants to our data: (1) the fixed model, in which a single attraction rate and a single leaving rate were estimated for both choices, representing a null hypothesis of no preference between choices; (2) the free attraction model, in which the attraction rates to each plant were free to vary but the leaving rates from each were set equal to each other; (3) the free leaving model, in which attraction rates were fixed but leaving rates were free to vary; and (4) the free choice model, in which both the attraction and leaving rates were free to vary. We compared model variants using Akaike’s information criterion corrected for small sample size (AICc) and calculated variances for each parameter using the quadratic approximation method (Bolker 2008). We calculated model-averaged parameter estimates and variances using the AICc weights of all models with values of ΔAICc ≤ 7 (Burnham et al. 2011) and calculated 95% confidence intervals (CI) from these model-averaged variances (Zeilinger et al. 2014). We inferred vector preference between host choices (infected source plant vs. test plant) using both model selection and 95% CI.

We analyzed the change in vector acquisition and transmission of *X. fastidiosa* over the weeks post-inoculation by fitting a series of nonlinear models to our data: two unimodal models—Holling type IV and Ricker models—and two saturating models—Holling type II and logistic growth models (Appendix S3: Table S1). We also fit data to a linear model as a null hypothesis to test against our nonlinear model set. We selected the best model using AICc as described in Appendix S3. Based on the best fitting model, we compared 95% CI of parameter estimates to make inferences on differences in the dynamics between resistant and susceptible genotypes.

Finally, we investigated which set of transmission-related parameters best explained variation in vector transmission, that is, infection status of test plants, using piecewise structural equation modeling (SEM). We were primarily interested in the paths through which the *PdR1* resistance trait influences transmission and in the relative importance of different components of host selection by vectors on transmission. We constructed a series of (generalized) linear models based on our hypothesized causal linkages or paths among the explanatory variables leading to infection of test plants. We evaluated the importance of the paths with model selection and standardized coefficient estimates (Appendix S4).

In discussing results of statistical tests below, we prefer the language of clarity over the use of significance, following the arguments of Dushoff et al. (2019). Also note that we rely more on 95% CI and AIC values rather than *P*-values in making inferences (Nakagawa and Cuthill 2007, Dushoff et al. 2019); nonetheless, we report *P*-values as well due to their prominence in traditional statistical reporting.

All analyses were performed in R 3.5.3 (R Core Team 2019). We used the rms package 5.1-3 to perform partial odds ordinal logistic regression (Harrell 2019). To fit consumer movement models, we used the optimx package version 7.10 with the spectral gradient optimization algorithm (Nash and Varadhan 2011); to fit nonlinear transmission models, we used the bbmle package 1.0.20 (Bolker and R Core Team 2017). For structural equation modeling, we used the piecewiseSEM package 2.0.2 (Lefcheck 2016). All data have been archived (Zeilinger et al. 2021), and all R code for analyses have been archived at https://doi.org/10.5281/zenodo.4547775.

**Results**

**Pierce’s disease severity and *X. fastidiosa* population dynamics**

In both years, inoculated susceptible grapevines exhibited Pierce’s disease symptoms earlier and more severely than inoculated resistant grapevines (Fig. 2A, B). In 2016, we could not detect a significant effect of week or difference in genotypes (weeks post-inoculation estimate ± standard error [SE] = 2.692 ± 6.885, Wald *Z* = 0.39, *P* = 0.696; genotype estimate ± SE = 1.677 ± 75.024, Wald *Z* = 0.02, *P* = 0.982; week-by-genotype interaction estimate ± SE = 1.137 ± 8.074, *Z* = 0.14, *P* = 0.888). In 2017, while we did not detect any clear difference in the genotype main effect, there was a clear trend of increasing symptom severity over time and a
Fig. 2. Mean ± standard error (SE) Pierce’s disease symptom severity index (A, B), mean ± SE population

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(Fig. 2. Continued)
sizes of *Xylella fastidiosa* in inoculated source plants (C, D), vector acquisition (E, F), and transmission of *X. fastidiosa* to test plants (G, H). Panels in the left column show results from 2016 (A, C, E, G); panels in the right column are from 2017 (B, D, F, H). Susceptible genotypes are represented by open symbols and dashed lines; *PdR1*-resistant genotypes are in closed symbols and solid lines. Lines in panels of vector acquisition and transmission (E–H) represent predictions from best nonlinear model based on Akaike’s information criterion corrected for small sample size. The legend in (B) indicates the different genotypes used in 2017; points in (F, H) represent data combined over genotypes. Population sizes are expressed as colony-forming units (CFU) g⁻¹ of fresh plant tissue and log₁₀ transformed, though untransformed data were used for statistical analysis (see Statistical analysis section). Error bars represent ± SE.

clear interaction between week and genotype (week estimate ± SE = 0.706 ± 0.141, Z = 5.00, P < 0.0001; genotype estimate ± SE = 1.526 ± 1.873, Z = 0.81, P = 0.415; week-by-genotype interaction estimate ± SE = −0.463 ± 0.183, Z = −2.53, P = 0.012).

Both susceptible and resistant grapevines harbored substantial populations of *X. fastidiosa* throughout our experiments, though populations in resistant lines were consistently lower (Fig. 2C, D). In 2016, *X. fastidiosa* populations increased over time and were clearly greater in susceptible compared with resistant lines (week post-inoculation coefficient estimate [95% CI] = 0.321 [0.0914, 0.660], t = 2.32, P = 0.025; genotype estimate [95% CI] = 2.96 [0.340, 6.94], t = 1.85, P = 0.072). There was no clear interaction between weeks and genotype (estimate [95% CI] = −0.122 [−0.471, 0.128], t = −0.835, P = 0.41). In 2017, we found no clear increase in *X. fastidiosa* populations over time (weeks post-inoculation estimate [95% CI] = 0.0382 [−0.0262, 0.102], t = 1.18, P = 0.24). However, *X. fastidiosa* populations were substantially lower in the resistant genotypes than the susceptible genotype 007, which was used as a baseline for statistical comparison (resistant line 094 estimate [95% CI] = −1.91 [−3.71, −0.448], t = −2.34, P = 0.021; resistant line 102 estimate [95% CI] = −1.86 [−3.83, −0.293], t = −2.11, P = 0.037).

**Host selection by vectors**

In the 2016 experiment, the fixed model had the lowest ΔAICc for each genotype at 3 and 8 weeks post-inoculation as well as for trials with the resistant genotype at 12 weeks (Table 1). The fixed model was clearly the best for most of these trials (i.e., the ΔAICc values for the remaining models were >2), indicating that the *G. atropunctata* showed no preference between inoculated source plants and uninfected test plants. The sole exception was in trials with susceptible source plants at 12 weeks post-inoculation; here, the free leaving model fits the data best and clearly much better than the fixed model (Table 1), indicating that *G. atropunctata* showed a preference between the host plant choices. However, the 95% CI for the attraction and leaving rates were quite large (Fig. 3A), muddling any apparent differences.

In the 2017 experiment, patterns of preference were more complicated than in 2016 though the differences were also clearer (Table 1). At two weeks post-inoculation, the fixed model fits the data best overall. While the fixed model performed slightly worse than the free attraction and free leaving models for trials with genotype 092 susceptible source plants, the difference was negligible (ΔAICc < 1). At five weeks post-inoculation, the fixed model fits the data best for trials with genotypes 007 susceptible and 102 resistant; once again, the fixed model performed worse than the free attraction and free leaving models for genotype 092 susceptible but now the difference in performance was stronger (ΔAICc > 2). Interestingly, for genotype 094 resistant, the fixed model performed much worse than all other models, with the free attraction model performing best (Table 1).

At eight weeks post-inoculation, the fixed model performed the best for both susceptible genotypes (007 and 092), whereas the free attraction model performed the best for both resistant genotypes (094 and 102; Table 1).
Table 1. Model selection tables for consumer movement model variants for each combination of week post-inoculation and genotype for the 2016 and 2017 transmission experiments.

| Genotype by no. weeks post-inoculation | Model     | AICc   | ΔAICc | df |
|---------------------------------------|-----------|--------|-------|----|
| 2016                                  |           |        |       |    |
| 3 weeks                               |           |        |       |    |
| Susceptible                           | Fixed     | 171.90 | 0     | 2  |
|                                       | Free leaving | 176.47 | 4.57  | 3  |
|                                       | Free attraction | 177.46 | 5.56  | 3  |
|                                       | Free choice   | 183.90 | 12.00 | 4  |
| Resistant                             | Fixed     | 146.47 | 0     | 2  |
|                                       | Free leaving | 152.01 | 5.54  | 3  |
|                                       | Free attraction | 152.06 | 5.59  | 3  |
|                                       | Free choice   | 161.34 | 14.87 | 4  |
| 8 weeks                               |           |        |       |    |
| Susceptible                           | Fixed     | 153.10 | 0     | 2  |
|                                       | Free leaving | 158.61 | 5.51  | 3  |
|                                       | Free attraction | 158.63 | 5.53  | 3  |
|                                       | Free choice   | 167.94 | 14.85 | 4  |
| Resistant                             | Fixed     | 132.33 | 0     | 2  |
|                                       | Free leaving | 137.22 | 4.89  | 3  |
|                                       | Free attraction | 137.84 | 5.51  | 3  |
|                                       | Free choice   | 146.54 | 14.22 | 4  |
| 12 weeks                              |           |        |       |    |
| Susceptible                           | Fixed     | 158.39 | 0     | 3  |
|                                       | Free attraction | 159.81 | 1.41  | 3  |
|                                       | Fixed     | 161.58 | 3.19  | 2  |
|                                       | Free choice   | 167.52 | 9.13  | 4  |
| Resistant                             | Fixed     | 145.83 | 0     | 2  |
|                                       | Free attraction | 148.75 | 2.92  | 3  |
|                                       | Free leaving | 149.49 | 3.66  | 3  |
|                                       | Free choice   | 157.95 | 12.12 | 4  |
| 2017                                  |           |        |       |    |
| 2 weeks                               |           |        |       |    |
| 007 Susceptible                       | Fixed     | 138.32 | 0     | 2  |
|                                       | Free leaving | 143.53 | 5.21  | 3  |
|                                       | Free attraction | 143.59 | 5.27  | 3  |
|                                       | Free choice   | 152.85 | 14.52 | 4  |
| 092 Susceptible                       | Free leaving | 127.55 | 0     | 3  |
|                                       | Free attraction | 127.83 | 0.28  | 3  |
|                                       | Fixed     | 128.43 | 0.88  | 2  |
|                                       | Free choice   | 136.37 | 8.82  | 4  |
| 094 Resistant                         | Fixed     | 137.87 | 0     | 2  |
|                                       | Free leaving | 140.58 | 2.71  | 3  |
|                                       | Free attraction | 140.60 | 2.72  | 3  |
|                                       | Free choice   | 149.65 | 11.78 | 4  |
| 102 Resistant                         | Fixed     | 134.87 | 0     | 2  |
|                                       | Free attraction | 137.72 | 2.84  | 3  |
|                                       | Free leaving | 138.44 | 3.57  | 3  |
|                                       | Free choice   | 147.05 | 12.17 | 4  |
| 5 weeks                               |           |        |       |    |
| 007 Susceptible                       | Fixed     | 141.89 | 0     | 2  |
|                                       | Free leaving | 147.42 | 5.53  | 3  |
|                                       | Free attraction | 147.47 | 5.57  | 3  |
|                                       | Free choice   | 156.70 | 14.80 | 4  |
At 14 weeks post-inoculation, the free attraction model performed the best for all genotypes. At the same time, the performance of the free attraction model as the best model was much clearer for the susceptible genotypes than for the resistant genotypes (Table 1).

Notes: Separate sub-tables are sorted by ΔAICc in ascending order. Abbreviations are as follows: R, resistant genotype; S, susceptible genotype; AICc, Akaike’s information criterion corrected for small sample size; ΔAICc, difference between lowest AICc score and a given score; df, degrees of freedom, that is, number of model parameters.
Comparing across years, overall movement rates were not clearly greater in one year over the other. Attraction rates were on average greater in 2016 \( (2016 \text{ mean attraction rate} = 0.713 \text{ h}^{-1}; \ 2017 \text{ mean attraction rate} = 0.633 \text{ h}^{-1}) \) while leaving rates were much greater in 2017 \( (2016 \text{ mean leaving rate} = 0.0165 \text{ h}^{-1}; \ 2017 \text{ mean leaving rate} = 0.0673 \text{ h}^{-1}) \).
Vector acquisition and transmission

For the vector acquisition analysis, unimodal models performed the best in both years. In 2016, the Ricker model fits the data best for both the susceptible and resistant genotypes (Table 2). While vector acquisition was consistently greater from susceptible grapevines, the parameter estimates did not differ based on 95% CI (Table 3, Fig. 1E). In 2017, the Holling type IV model fits the acquisition data best for both susceptible and resistant genotypes (Table 2). In this case, 95% CI did not overlap for two out of three of the parameters estimated; CI could not be estimated for the third parameter (Table 3). Acquisition rates were clearly greater from susceptible plants than resistant genotypes (Fig. 1F).

Vector transmission (i.e., the proportion of test plants infected with X. fastidiosa) in 2016 was overall slightly greater from susceptible plants (total proportion of test plants infected = 0.538) than from resistant plants (0.435). AICc values among models were relatively similar (Table 3); three models fit the data from the resistant genotype equally well: the logistic growth, Ricker, and linear models. However, from the parameter estimates, the fits of the logistic growth and linear models resulted in very high transmission rates at the y-intercept, that is, immediately at the time of inoculation, which would be biologically implausible. We selected the Ricker model as it predicts a transmission rate near zero at the y-intercept. Likewise, the logistic growth, Ricker, and linear models fit the data from the susceptible genotype in 2016 similarly, though the Ricker model was slightly better in this case and again more biologically plausible. The initial increase in transmission rate was similar between the two genotypes—as seen in similar estimates for parameter a (Table 3)—but the timing of peak transmission differed slightly—with 95% CI of the parameter b only slightly overlapping (Table 3); transmission from resistant genotypes peaked much earlier than from susceptible genotypes (Fig. 2G). Both curves suggest a similar probability of transmission at the peak (~0.75).

In 2017, transmission was overall slightly greater from resistant plants (total proportion of test plants infected = 0.234) than from

### Table 2. Results of model selection among nonlinear models for vector acquisition and transmission in 2016 and 2017.

| Genotype by year and response | Model         | AICc | ΔAICc | df |
|------------------------------|---------------|------|-------|----|
| Acquisition                  |               |      |       |    |
| 2016                         |               |      |       |    |
| Resistant                    | Ricker        | 15.49| 0      | 2  |
| Holling type IV              | 17.71         | 2.21 | 3      |    |
| Logistic growth              | 18.87         | 3.38 | 2      |    |
| Linear                       | 19.09         | 3.59 | 2      |    |
| Holling type II              | 20.07         | 4.58 | 2      |    |
| Susceptible                  | Ricker        | 15.41| 0      | 2  |
| Holling type IV              | 17.63         | 2.21 | 3      |    |
| Holling type II              | 22.95         | 7.54 | 2      |    |
| Logistic growth              | 23.70         | 8.29 | 2      |    |
| Linear                       | 23.71         | 8.29 | 2      |    |
| 2017                         |               |      |       |    |
| Resistant                    | Holling type IV| 32.13| 0      | 3  |
| Ricker                       | 56.70         | 24.58| 2      |    |
| Holling type II              | 62.26         | 30.13| 2      |    |
| Logistic growth              | 71.56         | 39.44| 2      |    |
| Linear                       | 74.55         | 42.43| 2      |    |
| Susceptible                  | Holling type IV| 59.73| 0      | 3  |
| Ricker                       | 85.56         | 25.83| 2      |    |
| Holling type II              | 94.61         | 34.88| 2      |    |
| Logistic growth              | 109.31        | 49.58| 2      |    |
| Linear                       | 112.50        | 52.77| 2      |    |
| Transmission                 |               |      |       |    |
| 2016                         |               |      |       |    |
| Resistant                    | Logistic growth| 13.14| 0      | 2  |
| Ricker                       | 13.15         | 0.01 | 2      |    |
| Linear                       | 13.16         | 0.01 | 2      |    |
| Holling type II              | 14.21         | 1.07 | 2      |    |
| Susceptible                  | Logistic growth| 18.74| 5.59  | 3  |
| Ricker                       | 14.34         | 0    | 2      |    |
| Linear                       | 15.35         | 1.01 | 2      |    |
| Holling type II              | 15.4          | 1.06 | 2      |    |
| Holling type IV              | 16            | 1.66 | 2      |    |
| Susceptible                  | Holling type IV| 19.4 | 5.06  | 3  |
| 2017                         |               |      |       |    |
| Resistant                    | Holling type IV| 16.95| 0      | 3  |
| Linear                       | 19.87         | 2.92 | 2      |    |
| Ricker                       | 19.88         | 2.93 | 2      |    |
| Holling type II              | 20.06         | 3.11 | 2      |    |
| Logistic growth              | 22.55         | 5.6  | 2      |    |
| Susceptible                  | Ricker        | 16.37| 0      | 2  |
| Holling type II              | 16.97         | 0.6  | 2      |    |
| Linear                       | 17.93         | 1.56 | 2      |    |
| Logistic growth              | 18.05         | 1.68 | 2      |    |
| Holling type IV              | 18.7          | 2.33 | 3      |    |

Notes: Separate sub-tables are sorted by ΔAICc in ascending order. Abbreviations as in Table 1.
susceptible plants (0.203), though transmission from resistant plants was much more dynamic (Fig. 2H). The Ricker model fit the data from susceptible genotypes the best, though the Holling type II provided a similar fit (Table 3). In contrast, the Holling type IV model fits the data from resistant genotypes the best. That different models were selected indicates qualitatively different dynamics for transmission from the resistant and susceptible genotypes. Comparing transmission between years, the peak transmission in 2017 was later and lower than in 2016 for both genotypes (Fig. 2G, H). Model selection results and parameter estimates are discussed in more detail in Appendix S3.

Finally, we synthesized all our data from 2017 on the vector transmission process into a structural equation model (as pointed out in the Materials and Methods, we lacked sufficient sample size in 2016). The best structural equation model—based on AICc values—included all variables except Pierce’s disease symptom severity index, although the model excluding leaving rates as well was nearly indistinguishable (Appendix S4: Table S1). The absolute fit of the best model—based on Fisher’s C statistic—was appropriate given the data (Fisher’s C = 32.83, df = 32, P = 0.43). Variation in transmission rate was best explained by variation in the vector acquisition rate, which in turn was best explained by X. fastidiosa population size in source plants (Fig. 4). Population size itself was best explained by the presence or absence of the PdR1 resistance trait. All components of vector preference remained in the best model, with leaving rates more strongly influencing vector transmission than attraction rates.

DISCUSSION

Vector-borne plant pathogens threaten agricultural sustainability, and development of disease-resistant cultivars for agricultural species provides an important management alternative to

| Genotype by year and response | Model          | Parameter | Estimate [2.5%, 97.5% CI] |
|------------------------------|----------------|-----------|---------------------------|
| Acquisition                  |                |           |                           |
| 2016                         | Resistant      | Ricker    | a 0.366 [0.238, 0.508]    |
|                              |                |           | b 0.16 [0.139, 0.2]       |
| Susceptible                  | Ricker         | a 0.394 [0.315, 0.466] |
| 2017                         |                | b 0.152 [0.145, 0.175]  |
| Resistant                    | Holling type IV| a 0.049 [0.04, 0.057]| |
|                              |                | b 50.364 [ND, ND]      |
| Susceptible                  | Holling type IV| a 0.111 [0.099, 0.123]|
|                              |                | b 52.354 [ND, ND]      |
| Transmission                 |                | c −13.35 [−13.353, −13.348]|
| 2016                         | Resistant      | Ricker    | a 0.704 [0.291, 1.032]    |
|                              |                | b 0.344 [0.27, 0.524]  |
| Susceptible                  | Ricker         | a 0.399 [0.172, 0.575] |
|                              |                | b 0.196 [0.151, 0.294] |
| 2017                         | Resistant      | Holling type IV| a 0.069 [ND, ND]    |
|                              |                | b 62.221 [ND, ND]      |
| Susceptible                  | Holling type IV| c −14.517 [ND, ND]    |
|                              | Ricker         | a 0.07 [0.021, 0.189]  |
|                              |                | b 0.111 [0.011, 0.238] |

Note: ND, not determined; CI could not be calculated by inverting the Hessian matrix.
insecticides. At the same time, partially resistant or tolerant cultivars could—according to theoretical predictions—increase the risk of disease spread and spillover to more susceptible hosts (Zeilinger and Daugherty 2014, Sisterson and Stenger 2018). In the context of such epidemic risks, we tested the impacts of the PdR1 resistance trait in grapevines on transmission of the bacterial pathogen X. fastidiosa by the insect vector G. atropunctata. Our results confirmed previous findings that PdR1 grapevines exhibit reduced disease severity and harbor lower X. fastidiosa population sizes than susceptible genotypes (Krivanek et al. 2005, Krivanek and Walker 2005, Fritschi et al. 2007).

Host defense against pathogens—whether resistance or tolerance—is predicted to have widely divergent epidemic consequences depending on how host-selection behavior of vectors is influenced by host infection: Resistance should generally reduce pathogen spread whereas tolerance traits combined with avoidance of diseased hosts should enhance pathogen spread (Zeilinger and Daugherty 2014). In our 2017 experiments, we found evidence of avoidance by G. atropunctata of infected host plants only when disease symptoms were severe—at late stages of infection (14 weeks post-inoculation) in susceptible plants. Avoidance of severely diseased susceptible plants of the 007 genotype was clear from both model selection by AICc and model-averaged 95% CI around attraction rate estimates (Table 1, Fig. 3B). For the 092 susceptible genotype, the 95% CI overlapped while model selection results strongly suggested preference, making evidence of avoidance less clear. In 2016, the results of our model selection process suggest avoidance of diseased susceptible plants; however, the wide CI make inferences still more difficult (Table 1, Fig. 3A). Previous work also found that cicadellid vectors of X. fastidiosa—including our study species—consistently avoid symptomatic infected host plants (Marucci et al. 2005, Daugherty et al. 2011, 2017, De Miranda et al. 2013). Furthermore, when avoidance was clear in our study, it was realized through differences in attraction rates, corroborating previous research that cicadellid vectors orient toward
host plants using visual cues (Rashed et al. 2011). Unexpectedly, we also found evidence of preference toward infected resistant plants at intermediate stages of infection (8 weeks post-inoculation). As no previous studies report similar findings, it remains unclear what may be driving this preference.

Vector acquisition rates—measured as the proportion of vectors that became infectious—were clearly nonlinear over the course of disease progression, following unimodal dynamics. Furthermore, vector transmission rates—measured as the proportion of healthy test plants that became infected—exhibited similar unimodal dynamics. Both acquisition and transmission rates were consistently greater in 2016 than 2017, likely because we ran trials for twice as long in 2016.

Contrary to predictions from theory (Fig. 1; Daugherty et al. 2017), we saw similar dynamics in transmission from both susceptible and PdR1-resistant grapevines. For transmission from susceptible vines, our results broadly conform to our predictions of unimodal transmission dynamics: transmission increases early after inoculation coinciding with increasing population size of X. fastidiosa, then declines with increasing symptom severity and concomitant avoidance by vectors. For trials with resistant vines however, because of a lack of symptom development, the unimodal dynamics appear to be more strongly tied to X. fastidiosa population size; specifically, transmission appears to decline in later stages of infection because of substantial declines in X. fastidiosa population size in two of the three resistant genotypes that we tested. Thus, the available evidence suggests distinct mechanisms underlying these similar patterns. In 2017, different models fit our transmission data from the different genotypes—with the Ricker model for susceptible plants and the Holling type IV model for resistant plants—suggesting qualitatively different dynamics and corroborating our interpretation. Previous studies of PdR1 resistance report similar findings. Fritschi et al. (2007) found unimodal X. fastidiosa population dynamics in a range of resistant Vitis spp., including an accession of V. arizonica related to the parental background of the grapevines used in our study; they suggested that such unimodal dynamics could be caused by induced resistance. This was echoed by Riaz et al. (2018) who hypothesized that PdR1 could confer a broad set of constitutive and inducible defensive traits against pathogens.

Beginning with the seminal work of McElhany et al. (1995), theory on the epidemiology of vector-borne plant diseases has increasingly emphasized the importance infection-induced host-selection behavior by vectors. Meanwhile, experimental work has succeeded in documenting a wide variety of infection-induced host-selection behaviors in vectors (reviewed in Eigenbrode et al. 2018). However, the theory remains disconnected from the empirical work because the vast majority of the empirical studies fail to test whether vector preference has any effect on transmission of pathogens to new hosts. While a handful of studies have suggested a positive or neutral association between infection-induced vector preference and transmission (Jennersten 1988, Daugherty et al. 2011, 2017, Del Cid et al. 2018), no study has yet to quantitatively test for such an empirical relationship.

We were able to efficiently determine the relationships between vector preference and transmission by decomposing preference into attraction and leaving rates using the consumer movement model of Gray et al. (2020) and incorporating these estimates into a structural equation model (SEM) of the vector transmission process. We found that while host selection by vectors is important for determining transmission, these components were far less important than vector acquisition rate, which was mostly driven by population size of X. fastidiosa in the infected plants, which in turn was determined by the presence or absence of the PdR1 resistance trait. Furthermore, leaving rates from infected as well as healthy plants tended to be stronger predictors of transmission than their respective attraction rates.

Our results provide only partial support to the existing theory. Attraction rates and leaving rates of G. atropunctata undoubtedly influenced transmission dynamics of X. fastidiosa to some degree; we observed reduced attraction rates toward susceptible plants at later stages of disease that clearly coincided with—and provide the only plausible explanation for—a decline in transmission. However, our SEM analysis contradicts the results of Sisterson (2008) and Shaw et al. (2017) who predict a much stronger influence of
attraction rates over leaving rates in driving pathogen spread. At the same time, the apparent importance of the leaving rate from healthy hosts supports the theory of Madden et al. (2000), who predicted that transmission of persistent pathogens was sensitive to tenure time on healthy predicted that transmission of persistent pathogens was sensitive to tenure time on healthy hosts (where tenure time = leaving rate $^{-1}$).

Much of the theory on vector-borne plant diseases was developed in the context of the host manipulation hypothesis and as such assumes that infection-induced vector preference is adaptive for the pathogen (Mauck et al. 2018). In our system, vector avoidance of diseased plants is unlikely to be adaptive for X. fastidiosa, an important distinction from the systems motivating much of the theory. What’s more, the effect size of preference in our study was relatively modest and inconsistent. Nonetheless, the theory should still apply to all cases of infection-induced vector preference; the relationship between preference and spread has largely been modeled using monotonic relationships, implying that the relationship should remain regardless of the direction and magnitude of preference (McElhany et al. 1995, Sisterson 2008, Zeilinger and Daugherty 2014, Shaw et al. 2017). For example, Shaw et al. (2017) predict that avoidance of infected hosts—such as what we find in our system—should broadly reduce spread ($\delta$ and $\varepsilon$ in their model). However, we failed to find such a clear relationship. At the same time, the over-riding importance of vector acquisition in our results provides some indirect support for the prediction of Shaw et al. (2017) that abundance of infectious vectors is more influential than vector behavior.

What could lead to a disconnect between host-selection behavior of vectors and pathogen transmission? If a vector prefers feeding on an infected plant, why would this not inevitably lead to greater transmission or, inversely, why would avoidance not inevitably lead to reduced transmission? One possibility is that herbivorous insects suffer from well-documented neurological limitations in processing sensory inputs, leading them to not always make the apparently best host selections (Bernays 2001). These neural limitations are hypothesized to be greater for generalist herbivores. Thus, with more complex environmental stimuli, we might expect the relationship between infection-induced vector preference and transmission to break down further, for example, when going from greenhouse experiments to field conditions. Such hypotheses could be incorporated into the theory with stochastic models, though only Donnelly et al. (2019) have so far considered stochasticity in their modeling. Clearly, further empirical work that test the relationship between vector preference (or manipulation) and transmission are needed to properly assess the generality of current theory across plant disease systems.

While the use of resistant cultivars is a key component to integrated disease management in agriculture, the risks of disease spillover or epidemics from such cultivars should be assessed prior to release. Generally, the greatest risk of transmission of vector-borne plant pathogens should occur within a window of disease progression when both pathogen burden and attractiveness for the vector are high, all else being equal (De Moraes et al. 2014). These two processes are clearly dynamic and may be governed by different drivers. For disease systems where vectors avoid symptomatic hosts, such as X. fastidiosa-associated diseases, the post-latent incubation period—in which hosts are infectious but asymptomatic—should be the period of greatest vector acquisition and thus most critical for pathogen spread (Daugherty et al. 2017, Kyrkou et al. 2018, Bragard et al. 2019).

Our results indicate that $PdR1$ hybrid grapevines can produce transmission rates greater than those from susceptible vines, potentially posing a risk of enhanced spread of X. fastidiosa. However, our results also suggest that these higher transmission rates are transient, followed by transmission rates similar to or lower than those from susceptible plants. Importantly, our structural equation modeling indicates that the $PdR1$ trait influences vector transmission much more strongly through X. fastidiosa population size than through the progression of disease symptoms. This suggests that the $PdR1$ trait operates more like resistance than tolerance. Further work is needed to examine additional environmental factors that could contribute to shortening or broadening the incubation period, such as temperature and water stress (Daugherty et al. 2017, Del Cid et al. 2018). Validating the dynamics explored here under field conditions will also be critical. We might expect transmission dynamics in the field to proceed over a longer time scale with larger
more established vines—potentially over multiple seasons—and operate with greater stochasticity due to greater environmental variation.

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LITERATURE CITED

Anderson, P. K., A. A. Cunningham, N. G. Patel, F. J. Morales, P. R. Epstein, and P. Daszak. 2004. Emerging infectious diseases of plants: pathogen pollution, climate change and agrotechnology drivers. Trends in Ecology & Evolution 19:535–544.

Bernays, E. A. 2001. Neutral limitations in phytophagous insects: implications for diet breadth and evolution of host affiliation. Annual Review of Entomology 46:703–727.

Best, A., A. White, and M. Boots. 2008. Maintenance of host variation in tolerance to pathogens and parasites. Proceedings of the National Academy of Sciences USA 105:20786–20791.

Blua, M. J., and T. M. Perring. 1992. Effects of Zucchini yellow mosaic virus on colonization and feeding behavior of *Aphis gossypii* (Homoptera: Aphididae) alatae. Environmental Entomology 21:578–585.

Bolker, B. M. 2008. Ecological models and data in R. Princeton University Press, Princeton, New Jersey, USA.

Bolker, B., and R Core Team. 2017. bbmle: tools for general maximum likelihood estimation. R package version 1.0.20. http://CRAN.R-project.org/package=bbmle

Borer, E. T., A.-L. Laine, and E. W. Seabloom. 2016. A multiscale approach to plant disease using the metacommunity concept. Annual Review of Phytopathology 54:397–418.

Bragard, C., et al. 2019. Update of the scientific opinion on the risks to plant health posed by *Xylella fastidiosa* in the EU territory. EFSA Journal 17:e05665.

Brown, J. K. M., and A. Tellier. 2011. Plant-parasite coevolution: bridging the gap between genetics and ecology. Annual Review of Phytopathology 49:345–367.

Burnham, K. P., D. R. Anderson, and K. P. Huyvaert. 2011. AIC model selection and multimodel inference in behavioral ecology: some background, observations, and comparisons. Behavioral Ecology and Sociobiology 65:23–35.

Chagnon, M., D. Kreutzweiser, E. A. D. Mitchell, C. A. Morrissey, D. A. Noome, and J. P. Van der Sluijs. 2015. Risks of large-scale use of systemic insecticides to ecosystem functioning and services. Environmental Science and Pollution Research International 22:119–134.

Cheatham, M. R., M. N. Rouse, P. D. Esker, S. Ignacio, W. Pradel, R. Raymundo, A. H. Sparks, G. A. Forbes, T. R. Gordon, and K. A. Garrett. 2009. Beyond yield: plant disease in the context of ecosystem services. Phytopathology 99:1228–1236.

Cooper, J. L., and A. T. Jones. 1983. Responses of plants to viruses: proposals for the use of terms. Phytopathology 73:127–128.

Daugherty, M. P., A. Rashed, R. P. P. Almeida, and T. M. Perring. 2011. Vector preference for hosts differing in infection status: sharpshooter movement and *Xylella fastidiosa* transmission. Ecological Entomology 36:654–662.

Daugherty, M. P., A. R. Zeilinger, and R. P. P. Almeida. 2017. Conflicting effects of climate and vector behavior on the spread of a plant pathogen. Phyto- biomes 1:46–53.

De Miranda, M. P., E. S. Villada, S. A. Lopes, A. Ferreres, and J. R. S. Lopes. 2013. Influence of citrus plants infected with *Xylella fastidiosa* on stilet penetration activities of *Bucephalogonia xanthiophis* (Hemiptera: Cicadellidae). Annals of the Entomological Society of America 106:610–618.

De Moraes, C. M., N. M. Stanczyk, H. S. Betz, H. Pulido, D. G. Sim, A. F. Read, and M. C. Mescher. 2014. Malaria-induced changes in host odors enhance mosquito attraction. Proceedings of the National Academy of Sciences USA 111:11079–11084.

Del Cid, C., R. Krugner, A. R. Zeilinger, M. P. Daugherty, and R. P. P. Almeida. 2018. Plant water stress and vector feeding preference mediate transmission efficiency of a plant pathogen. Environmental Entomology 46:1471–1478.
Donnelly, R., N. J. Cunniffe, J. P. Carr, and C. A. Gilligan. 2019. Pathogenic modification of plants enhances long-distance dispersal of non-persistently transmitted viruses to new hosts. Ecology 100:e02725.

Dushoff, J., M. P. Kain, and B. M. Bolker. 2019. I can see clearly now: reinterpreting statistical significance. Methods in Ecology and Evolution 10:756–759.

Dwyer, G., J. S. Elkinton, and J. P. Buonaccorsi. 1997. Xylella fastidiosa population dynamics in grapevine genotypes differing in susceptibility to Pierce’s disease. American Journal of Enology and Viticulture 48:868.

Eigenbrode, S. D., N. A. Bosque-Pérez, and T. S. Davis. 2018. Insect-borne plant pathogens and their vectors: ecology, evolution, and complex interactions. Annual Review of Entomology 63:169–191.

European Food Safety Authority (EFSA). 2018. Update of the Xylella spp. host plant database. EFSA Journal 16:e05408.

Fritschi, F. B., H. Lin, and M. A. Walker. 2007. Xylella fastidiosa population dynamics in grapevine genotypes differing in susceptibility to Pierce’s disease. American Journal of Enology and Viticulture 58:326–332.

Gilligan, C. A. 2008. Sustainable agriculture and plant diseases: an epidemiological perspective. Philosophical Transactions of the Royal Society B: Biological Sciences 363:741–759.

Gray, H., D. A. Andow, and K. Kiritani. 2020. Investigation of the movement components of host preference in a highly mobile insect herbivore, Nephotettix cincticeps (Hemiptera: Cicadellidae). Environmental Entomology 49:115–122.

Guilhabert, M. R., and B. C. Kirkpatrick. 2005. Identification of Xylella fastidiosa antivirulence genes: Hemagglutinin adhesins contribute to X. fastidiosa biofilm maturation and colonization and attenuate virulence. Molecular Plant-Microbe Interactions 18:856–868.

Harrell, F. 2019. rms: regression modeling strategies. R package version 5.1-3. https://CRAN.R-project.org/package=rms

Hill, B. L., and A. H. Purcell. 1995. Acquisition and retention of Xylella fastidiosa by an efficient vector, Graphocephala atropunctata. Phytopathology 85:209–212.

Hill, B. L., and A. H. Purcell. 1997. Populations of Xylella fastidiosa in plants required for transmission by an efficient vector. Phytopathology 87:1197–1201.

Hulcr, J., and R. R. Dunn. 2011. The sudden emergence of pathogenicity in insect–fungus symbioses threatens naive forest ecosystems. Proceedings of the Royal Society B: Biological Sciences 278:2866–2873.

Jennersten, O. 1988. Insect dispersal of fungal disease: effects of Ustilago infection on pollinator attraction in Viscaria vulgaris. Oikos 51:163–170.

Krivanek, A. F., S. Riaz, and M. A. Walker. 2006. Identification and molecular mapping of PdR1, a primary resistance gene to Pierce’s disease in Vitis. Theoretical and Applied Genetics 112:1125–1131.

Krivanek, A. F., J. F. Stevenson, and M. A. Walker. 2005. Development and comparison of symptom indices for quantifying grapevine resistance to Pierce’s disease. Phytopathology 95:36–43.

Krivanek, A. F., and M. A. Walker. 2005. Vitis resistance to Pierce’s disease is characterized by differential Xylella fastidiosa populations in stems and leaves. Phytopathology 95:44–52.

Kyrkou, I., T. Pusa, L. Ellegaard-Jensen, M.-F. Sagot, and L. H. Hansen. 2018. Pierce’s disease of grapevines: a review of control strategies and an outline of an epidemiological model. Frontiers in Microbiology 9:1–23.

Laine, A.-L., J. J. Burdon, P. N. Dodds, and P. H. Thrall. 2011. Spatial variation in disease resistance: from molecules to metapopulations. Journal of Ecology 99:96–112.

Lefcheck, J. S. 2016. piecewiseSEM: piecewise structural equation modeling in R for ecology, evolution, and systematics. Methods in Ecology and Evolution 7:573–579.

Legarrea, S., A. Barman, W. Marchant, S. Diffie, and R. Srinivasan. 2015. Temporal effects of a Begomovirus infection and host plant resistance on the preference and development of an insect vector, Bemisia tabaci, and implications for epidemics. PLOS ONE 10:e0142114.

Lu, G., T. Zhang, Y. He, and G. Zhou. 2016. Virus altered rice attractiveness to planthoppers is mediated by volatiles and related to virus titre and expression of defence and volatile-biosynthesis genes. Scientific Reports 6:38581.

Madden, L. V., M. J. Jeger, and F. van den Bosch. 2000. A theoretical assessment of the effects of vector-virus transmission mechanism on plant virus disease epidemics. Phytopathology 90:576–594.

Marucci, R. C., J. R. S. Lopes, J. D. Vendramim, and J. E. Corrente. 2005. Influence of Xylella fastidiosa infection of citrus on host selection by leafhopper vectors. Entomologia Experimentalis et Applicata 113:95–103.

Mauck, K. E., Q. Chesnais, and L. R. Shapiro. 2018. Evolutionary determinants of host and vector manipulation by plant viruses. Pages 189–250 in C. M. Malmstrom, editor. Environmental virology and virus ecology. Elsevier, Cambridge, Massachusetts, USA.
McElhany, P., L. A. Real, and A. G. Power. 1995. Vector preference and disease dynamics: a study of barley yellow dwarf virus. Ecology 76:444–457.

Miller, M. R., A. White, and M. Boots. 2006. The evolution of parasites in response to tolerance in their hosts: the good, the bad, and apparent commensalism. Evolution 60:945–956.

Nakagawa, S., and I. C. Cuthill. 2007. Effect size, confidence interval and statistical significance: a practical guide for biologists. Biological Reviews 82:591–605.

Nash, J. C., and R. Varadhan. 2011. Unifying optimization algorithms to aid software system users: optimx for R. Journal of Statistical Software 43:1–14.

R Core Team. 2019. R: a language and environment for statistical computing. R Foundation for Statistical Computing, Vienna, Austria.

Rashed, A., N. Killiny, J. Kwan, and R. P. P. Almeida. 2011. Background matching behaviour and pathogen acquisition: Plant site preference does not predict the bacterial acquisition efficiency of vectors. Arthropod-Plant Interactions 5:97–106.

Riaz, S., K. Huerta-Acosta, A. C. Tenscher, and M. A. Walker. 2018. Genetic characterization of Vitis germplasm collected from the southwestern US and Mexico to expedite Pierce’s disease-resistance breeding. Theoretical and Applied Genetics 131:1589–1602.

Roosien, B. K., R. Gomulkiewicz, L. L. Ingwell, N. A. Bosque-Pérez, D. Rajabaskar, and S. D. Eigenbrode. 2013. Conditional vector preference aids the spread of plant pathogens: results from a model. Environmental entomology 42:1299–1308.

Roy, B. A., and J. W. Kirchner. 2000. Evolutionary dynamics of pathogen resistance and tolerance. Evolution 54:51–63.

Savary, S., A. Ficke, J.-N. Aubertot, and C. Hollier. 2012. Crop losses due to diseases and their implications for global food production losses and food security. Food Security 4:519–537.

Shaw, A. K., A. Peace, A. G. Power, and N. A. Bosque-Pérez. 2017. Vector population growth and condition-dependent movement drive the spread of plant pathogens. Ecology 98:2145–2157.

Sicard, A., A. R. Zeilinger, M. Vanhove, T. E. Schartel, D. J. Beal, Daugherty, M. P., and R. P. P. Almeida. 2018. Xylella fastidiosa: insights into an emerging plant pathogen. Annual Review of Phytopathology 56:1–22.

Sisterson, M. S. 2008. Effects of insect-vector preference for healthy or infected plants on pathogen spread: insights from a model. Journal of Economic Entomology 101:1–8.

Sisterson, M. S., and D. C. Stenger. 2018. Modelling effects of vector acquisition threshold on disease progression in a perennial crop following deployment of a partially resistant variety. Plant Pathology 67:1388–1400.

Tumber, K., J. Alston, and K. Fuller. 2014. Pierce’s disease costs California $104 million per year. California Agriculture 68:20–29.

Walker, M. A., and A. C. Tenscher. 2016. Breeding Pierce’s disease resistant winegrapes. Pages 167–177 in Proceedings of the 2016 Pierce’s Disease Research Symposium. California Department of Food and Agriculture, San Diego, California, USA.

Werner, B. J., T. M. Mowery, N. A. Bosque-Pérez, H. Ding, and S. D. Eigenbrode. 2009. Changes in green peach aphid responses to potato leafroll virus—induced volatiles emitted during disease progression. Environmental Entomology 38:1429–1438.

Zeilinger, A. R., D. Beal, A. Sicard, C. M. Wallis, M. A. Walker, and R. P. P. Almeida. 2021. Pierce’s disease vector transmission-preference experiment on PdR1 resistant grapevines. Archived Data Sets. https://doi.org/10.5061/dryad.wstqjq2kf

Zeilinger, A. R., and M. P. Daugherty. 2014. Vector preference and host defense against infection interact to determine disease dynamics. Oikos 123:613–622.

Zeilinger, A. R., D. M. Olson, and D. A. Andow. 2014. A likelihood-based biostatistical model for analyzing consumer movement in simultaneous choice experiments. Environmental Entomology 43:977–988.

**Supporting Information**

Additional Supporting Information may be found online at: http://onlinelibrary.wiley.com/doi/10.1002/ecs2.3505/full