No Evidence for Orbital Clustering in the Extreme Trans-Neptunian Objects

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Abstract

The apparent clustering in longitude of perihelion \( \varpi \) and ascending node \( \Omega \) of extreme trans-Neptunian objects (ETNOs) has been attributed to the gravitational effects of an unseen 5–10 Earth-mass planet in the outer solar system. To investigate how selection bias may contribute to this clustering, we consider 14 ETNOs discovered by the Dark Energy Survey, the Outer Solar System Origins Survey, and the survey of Sheppard and Trujillo. Using each survey’s published pointing history, depth, and TNO tracking selections, we calculate the joint probability that these objects are consistent with an underlying parent population with uniform distributions in \( \varpi \) and \( \Omega \). We find that the mean scaled longitude of perihelion and orbital poles of the detected ETNOs are consistent with a uniform population at a level between 17% and 94% and thus conclude that this sample provides no evidence for angular clustering.

Unified Astronomy Thesaurus concepts: Solar system (1528); Planetary science (1255); Trans-Neptunian objects (1705); Kuiper belt (893); Detached objects (376)

1. Introduction

The apparent clustering in longitude of perihelion \( \varpi \) and ascending node \( \Omega \) of solar system bodies known as extreme trans-Neptunian objects (ETNOs) motivated the hypothesis that the solar system contains a 5–10 Earth-mass planet (Planet X/Planet 9) at 400–800 times Earth’s distance from the Sun (Trujillo & Sheppard 2014; Batygin & Brown 2016; Batygin et al. 2019). Some have proposed even more exotic sources of the apparent clustering, such as gravitational perturbations from a primordial black hole captured into orbit around the Sun (Scholtz & Unwin 2020).

While there is no universally accepted definition for the ETNOs, recent literature has emphasized objects with semimajor axis \( a \gtrsim 230 \) au and perihelion \( q > 30 \) au. Because ETNOs follow highly elliptical orbits, and their brightness decreases by \( 1/r^2 \), they are almost always discovered within a few decades of perihelion. Moreover, telescopic surveys observe a limited area of the sky, at particular times of the year, to a limited depth. These effects result in significant selection bias. The six ETNOs considered in the Batygin & Brown (2016, hereafter BB16) analysis were discovered in an assortment of surveys with unknown or unpublished selection functions, making it difficult to establish that the observed angular clustering was indeed of physical origin.

More recent surveys have carefully characterized their selection functions and applied these tools to small samples of new ETNOs. The Outer Solar System Origins Survey (OSSOS; Bannister et al. 2016) analyzed the bias present in the discovery of eight objects they detected with \( a > 150 \) au and \( q > 30 \) au (Shankman et al. 2017). They found that their detected objects were consistent with a uniform underlying population in \( \varpi \) and \( \Omega \). Bernardinelli et al. (2020a) analyzed samples of three to seven variously defined ETNOs discovered by the Dark Energy Survey (DES; DES Collaboration 2016; Dark Energy Survey Collaboration et al. 2016) and found the data consistent with angular isotropy.

Brown & Batygin (2019, hereafter BB19) attempted to reverse engineer the survey bias in the entire then-known population of 14 ETNOs using a sampling method (Brown 2017) on all TNOs known to the Minor Planet Center (MPC). In contrast to the individual survey-level analyses described above, BB19 concluded that the observed clustering is highly likely to be a physical effect, and they argued that the best explanation remains a massive distant planet.

While no single survey has discovered enough ETNOs to reach a statistically compelling conclusion, a stronger statement becomes possible when data from multiple surveys are combined. According to the criteria above, there are 14 ETNOs (Table 1) detected by three independent surveys with characterized selection functions, all published since BB16. Using the published pointing history, depth, and TNO tracking selections for DES (five objects; Khain et al. 2018; Bernardinelli et al. 2020b), OSSOS (five objects; Bannister et al. 2018), and the survey of Sheppard & Trujillo (2016, hereafter ST; four objects), we calculate the joint probability that these objects are consistent with the null hypothesis: an underlying population distributed uniformly in the longitudes \( \varpi \) and \( \Omega \). If the purported clustering is indeed a physical effect, we would expect it to remain consistent with the data in this larger, independent sample when selection functions are modeled.

2. Methods

The three surveys we consider have very different designs and scientific goals and, consequently, quite different ETNO selection functions. This is readily apparent from their survey footprints, shown in Figure 1. The DES, which was on-sky between 2012 and 2019, used the Dark Energy Camera (Flaugher et al. 2015) on the 4 m Blanco telescope at CTIO to carry out an extragalactic survey designed to measure cosmological parameters. It consisted of two interwoven surveys. In the 30 deg\(^2\) supernova survey, 10 separate fields were visited approximately weekly in the \( griz \) bands during the 6 months yr\(^{-1} \) that DES was in operation. In the 5000 deg\(^2\) wide survey, each field was imaged a total of 10 times at a sparse temporal cadence in each of the \( grizY \) bands over the duration of the survey. The wide survey reached a limiting \( r \)-band magnitude of \( \approx 23.5 \). The DES had limited near-ecliptic coverage centered near an ecliptic longitude of zero and a large off-ecliptic footprint that made it particularly sensitive to high-inclination objects. For our main analysis, we consider only the ETNOs detected in the DES wide survey and treat the supernova fields separately. The OSSOS survey (2013–2017), by contrast, was optimized to detect and track TNOs in eight \( \sim 20 \) deg\(^2\) blocks distributed along the ecliptic. This survey used the 3.6 m Canada–France–Hawaii Telescope and reached a limiting \( r \)-band magnitude of 24.1–25.2. Finally, the ST survey (2007–2015) used the Blanco, Subaru, Large Binocular, and Magellan telescopes to cover 1080 deg\(^2\) at an average distance of 135 from the ecliptic to a depth of approximately \( VR \sim 25 \). This survey aimed to detect the most distant objects: ETNOs and inner Oort cloud (IOC) objects such as Sedna. Therefore, only those candidates with an estimated heliocentric distance greater than 50 au were selected for follow-up and tracking.

The most complete way to account for survey bias in the discovery of the solar system objects is to use a survey simulator (Petit et al. 2011; Lawler et al. 2018). In essence, a survey simulator simulates detections of a model population of solar
system bodies by using a survey’s pointing history, depth, and tracking criteria. This allows for the computation of a survey’s selection function for a given population, which enables us to account for bias and therefore understand the true underlying populations. While it gives a reasonable approximation, the technique employed in BB19 cannot fully substitute for actually simulating each survey to calculate its selection function. Since the known ETNOs were discovered by a variety of surveys, the task of developing an appropriate simulator is nontrivial. Our simulator (FastSSim) is highly parametric, requiring only the few pieces of information common among all well-characterized surveys: pointing history, limiting magnitudes, and follow-up criteria. The basic flow of the simulator is as follows:

1. Map a survey’s published pointing history to a HEALPix grid, as in Figure 1.
2. Generate a distribution of fake objects at a single epoch.
3. Calculate the objects’ HEALPix pixels and apparent magnitudes.
4. Determine which fake objects fall in a survey’s footprint.
5. Make cuts according to the survey’s limiting magnitudes and follow-up criteria.

Note that this simulation method makes several approximations. We compute the sky coordinates of our objects at a single epoch, we use a single color and limiting magnitude for each survey field, we do not consider CCD-level detections (so we do not account for complications such as chip gaps), and we employ a step-function detection criterion (so we do not model survey cadence or linking efficiency). We use a single HEALPix pixel for each survey pointing. We have chosen the pixel scales for each telescope as follows: Blanco uses an NSIDE of 64 (except for the DES supernova fields, for which we use an NSIDE of 1024), and the Magellan, Large Binocular, and Subaru telescopes use an NSIDE of 128. These assumptions ignore the time history of the surveys, as well as the apparent motion of the objects. FastSSim works well for this

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**Figure 1.** HEALPix mapping of the currently released sky coverage of the three major TNO surveys of this generation. The surveys by OSSOS and ST hug the ecliptic plane (plotted in red), while DES, designed as a cosmological survey, has a much more expansive footprint.

**Table 1.** Barycentric Orbital Elements of the ETNOs Used in Our Analysis

| Object         | \(a\) (au) | \(e\)   | \(i\) (deg) | \(q\) (au) | \(\omega\) (deg) | \(\Omega\) (deg) | \(H\) (mag) | Survey        |
|----------------|------------|---------|-------------|------------|------------------|-----------------|------------|---------------|
| 2015 BP519    | 448.8      | 0.92    | 54.1        | 35.2       | 348.1            | 135.2           | 4.4        | DES           |
| 2013 SL102    | 314.3      | 0.88    | 6.5         | 38.1       | 265.5            | 94.7            | 7.1        | DES           |
| 2013 RA109    | 462.4      | 0.90    | 12.4        | 46.0       | 263.0            | 104.8           | 6.2        | DES           |
| 2014 WB56     | 289.1      | 0.85    | 24.2        | 42.5       | 234.6            | 114.9           | 7.3        | DES           |
| 2016 SG65     | 233.0      | 0.85    | 13.2        | 35.1       | 296.3            | 119.0           | 7.5        | DES           |
| 2013 SY40     | 733.1      | 0.93    | 4.2         | 50.1       | 32.2             | 29.5            | 6.7        | OSSOS         |
| 2015 RX345    | 426.4      | 0.89    | 12.1        | 45.7       | 65.1             | 8.6             | 6.2        | OSSOS         |
| 2015 GT10     | 311.4      | 0.88    | 8.8         | 38.5       | 129.0            | 46.1            | 8.5        | OSSOS         |
| 2015 KG163    | 679.7      | 0.94    | 14.0        | 40.5       | 32.1             | 219.1           | 8.2        | OSSOS         |
| uo5m93        | 283.0      | 0.86    | 6.8         | 39.5       | 43.3             | 165.9           | 8.8        | OSSOS         |
| 2013 FT28     | 295.4      | 0.85    | 17.4        | 43.4       | 40.7             | 217.7           | 6.7        | ST            |
| 2014 SR49     | 296.6      | 0.84    | 18.0        | 47.7       | 341.2            | 34.9            | 6.7        | ST            |
| 2015 TG387    | 1101.3     | 0.94    | 11.7        | 65.1       | 118.0            | 301.0           | 5.5        | ST            |
| 2014 FE72     | 1559.5     | 0.98    | 20.6        | 36.2       | 133.9            | 336.8           | 6.1        | ST            |
| 2012 VP113    | 262.7      | 0.69    | 24.1        | 80.5       | 293.8            | 90.8            | 4.0        | ST            |
| 2013 RF98     | 363.6      | 0.90    | 29.5        | 36.1       | 311.8            | 67.6            | 8.7        | DES SN        |

Note. All reported values are at the epoch JD 2,459,000.5 (except for uo5m93, whose elements are for the epoch JD 2,457,163.826). Here DES SN indicates discovery in the DES supernova fields. In order to maintain an independent sample from BB16, we do not include 2013 RF98 or 2012 VP113 in our main analysis. We discuss their effects separately.

46 The tools for the FastSSim algorithm have now been compiled into the open-source Python package SpaceRocks. It is under active development at https://github.com/knapier/spacerocks.
47 It is not important that we used a HEALPix mapping. We could have used any mapping onto the sphere.
application because the objects move slowly, the telescopes have large fields of view, and the sensitivity does not have much spatial variation.

We acquired the non-DES survey pointings and limiting magnitudes from ST, Sheppard et al. (2019), and Bannister et al. (2018). We choose each HEALPix pixel size to most closely match the field of view of the telescope used. This does not allow for a perfect mapping between pointings and pixels, but it turns out to be sufficient for our needs. In fact, we find that FastSSim performs remarkably well in cross-checks against both our full chip-level DES simulator (Hamilton 2019) and the OSSOS survey simulator described in Petit et al. (2011; see Figure 2). While FastSSim misses some of the fine details of the selection functions, the small sample of ETNOs and the approximate nature of this analysis make such fine details unimportant to our overall conclusion. Given the success of these cross-checks, we are confident in extending their use to characterize the survey of ST.

To simulate the surveys, we randomly generate ETNOs in accordance with a nominal scattered-disk model (specified in Section 3) until each survey has accumulated $10^5$ detections, to allow for high-resolution characterization of a survey’s sensitivity in $\varpi$ and $\Omega$. This typically requires the generation of approximately $10^{10}$ fakes, so our set of simulated objects spans the parameter space of the ETNOs. We consider an object to be detected if it is in one of the survey’s HEALPix pixels, is brighter than the pixel’s limiting magnitude, and has a perihelion distance $q > 30$ au. For the survey of ST, we satisfy a tracking criterion specified in Sheppard et al. (2016) by requiring an object to have a heliocentric distance of at least 50 au at the time of detection.

As a quantitative example of the effectiveness of FastSSim, Figure 2 shows a comparison with the CFEPs/OSSOS survey simulator. For this test, each simulator uses the population model defined in Section 3. Using Kuiper’s test, we find that the distributions of the 1615 data points calculated by FastSSim are statistically indistinguishable from those.
computed using the CFEPS/OSSOS survey simulator. Thus, in order to distinguish the two simulators, one would need ETNOs—well above the quantity discovered by OSSOS. We achieve similar results in quantitative comparisons against the distributions computed using the DES survey simulators (see Figure 4). Shankman et al. (2017) found a similar resilience to changes in pericenter distribution (see the Appendix for the distribution of orbital elements for populations with $q > 30, 35, \text{and } 38 \text{ au}$). Shankman et al. (2017) found a similar resilience to changes in the scattered-disk model. Noting the weak dependence of the outcome of our simulations on the choice of model, we proceed using the following scattered-disk model:

1. $a$ follows a single power-law distribution such that $N(a) \propto a^{0.7}$, where $a \in [230 \text{ au}, 1600 \text{ au}]$;
2. $e$ is distributed uniformly in $[0.69, 0.999]$;
3. $i$ follows a Brown distribution such that $N(i) \propto \sin(i) \exp \left[ -\frac{(i - \mu_i)^2}{2\sigma_i^2} \right]$, with $\mu_i = 0^\circ$ and $\sigma_i = 15^\circ$;
4. $H$ follows a single power-law distribution such that $N(H) \propto 10^{0.8H}$, where $H \in [4, 10]$; and
5. the perihelion distance $q > 30 \text{ au}$.

These model parameters produce posteriors in $a, e, q, i, H$ that appear to be in reasonable agreement with the real ETNO detections by each survey. See the Appendix for histograms of the posteriors in each of these variables, overlaid with a rug plot of each survey’s real detections.

### 3. Scattered-disk Model

To test the dependence of our analysis on the choice of scattered-disk model, we simulated models with various distributions of semimajor axis ($a$), eccentricity ($e$), inclination ($i$), and absolute magnitude ($H$) while keeping the orbital angles $\Omega$ and $\omega$ (and thus $\varpi$) uniform from $0^\circ$ to $360^\circ$. We tested manifold permutations with the parameter distributions: $N(a) \propto a^5$ with $a \in [230 \text{ au}, 1600 \text{ au}]$ and $\zeta \in [0.5, 1.0]$, uniform $i \in [0^\circ, 60^\circ]$, Brown distribution $i$ (Brown 2001) with a variety of widths ranging from $5^\circ$ to $25^\circ$, and $N(H) \propto 10^{0.5H}$ with $H \in [4, 10]$ and $\zeta \in [0.6, 0.9]$.

We found that our conclusions were not significantly affected by the variation of the model parameters. Our results are also robust to changes in pericenter distribution (see the Appendix for the distribution of orbital elements for populations with $q > 30, 35, \text{and } 38 \text{ au}$).

### 4. Analysis and Results

Performing a clustering analysis in the variables $\varpi$ and $\Omega$ is complicated, as the two are strongly correlated. We proceed by working in the orthogonal $\{x, y, p, q\}$ basis discussed in BB19.
importantly, \( \varpi \) and \( \Omega \) are linearly independent in this basis. Note that these vectors are not normalized but instead have their lengths modulated by eccentricity and inclination. The coordinates are defined as follows:

\[
G = \sqrt{1 - e^2} Z = \sqrt{1 - e^2} [1 - \cos(i)],
\]

\[
x = \sqrt{2G} \cos(\varpi) \quad y = \sqrt{2G} \sin(\varpi),
\]

\[
p = \sqrt{2Z} \cos(\Omega) \quad q = \sqrt{2Z} \sin(\Omega).
\]

Note that \( \Gamma \) and \( Z \) have been scaled by a factor of \( \sqrt{GM_\odot a} \) from their traditional forms, since the semimajor axis is not relevant to this argument. Figure 4 shows our calculated selection functions in the \( xy \)- and \( pq \)-planes.

For the sake of comparison, we used the method presented in BB19 to test the consistency of each survey’s detected ETNOs with its selection function. We first perform \( 10^6 \) iterations, sampling from our simulated detections a set of objects whose cardinality is equal to that of the set of real ETNOs detected by the given survey. We then take the average \( \{x, y, p, q\} \) position of each sample and use these values to construct a four-dimensional histogram. We display a Gaussian kernel density estimation of these data in the \( xy \)- and \( pq \)-planes in Figure 5.

We perform a Gaussian kernel density estimation on our mean-sampled histograms to obtain a probability distribution function (PDF). Next, we draw \( N \) samples from our simulated data (where \( N \) is the number of ETNOs actually detected by the survey), find the mean \( \{x, y, p, q\} \) position, and evaluate our PDF at that position. We repeat this \( 10^5 \) times to construct a likelihood function. Next, we compute this value for the ETNOs actually discovered by the survey. To calculate the probability of a survey detecting the ETNOs it actually detected (as opposed to some other set of ETNOs), we find the fraction of the \( 10^5 \) sample likelihood values that the survey’s actual likelihood value exceeds. Rounded to the nearest 1%, this probability for each survey is as follows: \( P_{\text{DES}} \sim 0.06 \), \( P_{\text{OSSOS}} \sim 0.53 \), and \( P_{\text{ST}} \sim 0.59 \).

The joint probability of \( N \) surveys detecting objects with given probabilities (or some less likely set of values) can be calculated as the volume under the surface of a constant product of probabilities in the domain of the \( N \)-dimensional unit hypercube, given by

\[
P_{\text{joint}} = P \sum_{k=0}^{N} (-1)^k \frac{\log(P)^k}{k!},
\]

where \( P \equiv \prod_{k} P_k \). In our case, \( k \in \{\text{DES}, \text{OSSOS}, \text{ST}\} \). Using Equation (4), we calculate the joint probability to be 24%.

With such a small sample size, this work is sensitive to outliers and the definition of “ETNO” itself. The high-inclination object 2015 BP519 is among the most dynamically anomalous objects in the solar system (Becker et al. 2018), and we cannot discount the possibility that it is of a different dynamical origin than the other ETNOs. If we redo our analysis without 2015 BP519, \( P_{\text{DES}} \) increases to 84%, and thus \( P_{\text{joint}} \) increases to \( \sim 85\% \). The object 2014 FE72 has an extremely large semimajor axis—roughly four standard deviations above

![Figure 5. Kernel density estimates of the mean \((x, y)\) and \((p, q)\) position of \(10^6\) samples of ETNOs drawn from the PDFs shown in Figure 4. The number of objects in each sample corresponds to the number of ETNOs detected by the given survey. The contours represent the samples (the contours scale linearly, and darker contours are more densely populated), while the red dots represent the mean position of the ETNOs detected by each survey.](image-url)
the mean of the ETNOs considered in this work. Its large semimajor axis carries it deep into the IOC region, where interactions with galactic tides make its secular relationship with a putative Planet X/Planet 9 less certain. If we exclude 2014 FE72, $P_{ST}$ increases to 88%, and thus $P_{joint}$ increases to 31%. If we include 2012 VP113, $P_{ST}$ increases to 60%, and $P_{joint}$ remains 24%. We also address the fact that the clustering by a putative Planet X/Planet 9 should be more robust in the sample of ETNOs with $q > 40$ au, since these objects avoid strong perturbations by Neptune. If we restrict our ETNOs to these eight objects, $P_{joint}$ increases to 94%. Finally, we analyze the subset of objects that are either stable or metastable in the presence of the putative Planet X/Planet 9 (Batygin et al. 2019): 2015 TG387, 2013 SY99, 2015 RX245, 2014 SR349, 2012 VP113, 2013 RA109, and 2013 FT28. For this subset, $P_{joint} = 82\%$.

For the sake of completeness, we also use a more traditional sampling method to determine the significance of the clustering of ETNOs. We begin by performing a Gaussian kernel density estimate on each survey’s posterior distributions. We then perform $10^5$ iterations in which we randomly draw $N$ points from each survey’s posterior distribution (where $N$ is the number of ETNOs detected by the survey) and multiply each of the $N$ probabilities together to calculate a likelihood. Finally, we calculate the same metric for each survey’s actual detections and compare the value to the distribution of our samples. As before, the probability for each survey is the fraction of the $10^5$ sample likelihood values that the survey’s actual likelihood value exceeds. Rounded to the nearest 1%, the probability for each survey is as follows: $P_{DES} \sim 0.06$, $P_{OSSOS} \sim 0.41$, and $P_{ST} \sim 0.43$. The joint probability is thus 17%.

For a more physically intuitive representation of the survey bias, refer to Figure 6. Here the radial quantity represents the barycentric distance, and the azimuthal quantity is true longitude (the true anomaly + $\omega$). The edge of the black circle is at 30 au. The white regions represent the combined surveys’ sensitivity (brighter regions correspond to higher sensitivity), weighted by the number of real ETNO detections. The red dots represent the real ETNOs at the epoch of discovery. The observations are in good agreement with the combined selection function, qualitatively confirming the conclusions of our formal statistical analysis performed on canonical variables.

### 4.1. DES Supernova Fields

The ETNO 2013 RF08 was discovered in the deep DES supernova (DES SN) fields. Since the DES SN fields are so small, they suffer from severe selection bias. Additionally, since their observing cadence and depth ($\sim 24.5$ in the $r$ band) are significantly different from those of the wide survey, they
need to be treated independently. We generated 1829 simulated detections in the DES SN fields (since the fields are so small, it is computationally prohibitive to generate $10^5$ synthetic detections, as we do for DES, OSSOS, and ST) from the population model defined in Section 3. Figure 7 shows a kernel density estimate of the detections in $(x, y)$ space. We show the posteriors in $(a, e, i, H, \Omega, \varpi)$ in Figures 8–14 in the Appendix. In all parameters, 2013 RF98 appears to be a rather ordinary detection for the DES SN fields.

Since there is only one data point here, we can just numerically integrate to find $P_{\text{DESSN}} = 0.33$ (i.e., a $p$-value of 0.33). Treating DES SN as its own survey, we may use Equation (4) to calculate the four-survey joint probability to find $P_{\text{joint}} = 25\%$.

5. Discussion and Conclusions

We use quantified selection bias calculations on all ETNOs discovered by the three most productive ETNO surveys, each with a quite different survey strategy and selection function, to test the consistency of the ETNOs with a uniform underlying distribution. Given a joint probability between 17% and 94% (i.e., a $p$-value between 0.17 and 0.94), we conclude that the sample of ETNOs from well-characterized surveys is fully consistent with an underlying parent population with uniform distributions in the longitudes $\varpi$ and $\Omega$. Our result differs drastically from the corresponding value in BB19 of 0.2%. Closer inspection sheds some light on the apparent discrepancy. If we examine only the overlapping set of ETNOs used this work and BB19 (2015 BP$_{519}$, 2013 RF$_{98}$, 2013 SY$_{99}$, 2015 RX$_{245}$, 2015 GT$_{50}$, 2015 KG$_{163}$, 2013 FT$_{28}$, 2014 SR$_{349}$, and 2014 FE$_{72}$), $P_{\text{joint}}$ drops to <0.005. This indicates an expected issue: small number statistics are sensitive to fluctuations. For example, when BB19 performed their analysis, a small but important set of ETNOs had not yet been reported to the MPC. As a concrete demonstration of the importance of the omission of a few ETNOs from BB19, consider DES. Of the five ETNOs discovered by the DES wide survey, BB19 included only 2015 BP$_{519}$. From Figure 4, it is clear that this object lands in an extremely low-probability region. This drives down $P_{\text{joint}}$ and thus gives a satisfactory answer as to why the result of this work differs so significantly from that of BB19.

It is important to note that our work does not explicitly rule out Planet X/Planet 9; its dynamical effects are not yet well enough defined to falsify its existence with current data. This work also does not analyze whether some form of clustering could be consistent with the 14 ETNOs we consider. For example, the ETNOs could happen to be clustered precisely where current surveys have looked. In that case, a survey with coverage orthogonal to the regions shown in Figure 6 would find far fewer ETNOs than expected. Various realizations of Planet X/Planet 9 predict clustering of various widths, modalities, and libration amplitudes and frequencies; we do not test for consistency with any of these distributions. Instead, we have shown that, given the current set of ETNOs from well-characterized surveys, there is no evidence to rule out the null hypothesis. Increasing the sample of ETNOs with ongoing and future surveys with different selection functions, such as the Deep Ecliptic Exploration Project (Trilling et al. 2019) and the Legacy Survey of Space and Time at the Vera Rubin Observatory (Schwamb et al. 2018), will allow for more restrictive results. Despite other lines of indirect evidence for Planet X/Planet 9, in the absence of clear evidence for clustering of the ETNOs, the argument becomes much weaker. Future studies should consider other mechanisms capable of giving the outer solar system its observed structure while preserving a uniform distribution of ETNOs in the longitudes $\Omega$ and $\varpi$.

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Appendix

In Figures 8–14 we show the posterior distributions of the orbital elements of our simulated detections.

![Figure 8](image-url)  
Figure 8. Posterior pericenter distance distributions of simulated detections. The red triangles are the real ETNO detections by each survey. Note that DES has partially overlapping data points at $q \approx 35$ au.
Figure 9. Posterior semimajor axis distributions of simulated detections. The red triangles are the real ETNO detections by each survey. Note that ST has partially overlapping data points at $a \approx 296$ au. The gray, red, and blue histograms correspond to cuts with $q > 30$, 35, and 38 au, respectively.
Figure 10. Posterior eccentricity distributions of simulated detections. The red triangles are the real ETNO detections by each survey. Note that DES has overlapping data points at $e = 0.85$. The gray, red, and blue histograms correspond to cuts with $q > 30$, $35$, and $38$ au, respectively.
Figure 11. Posterior inclination distributions of simulated detections. The red triangles represent the real ETNO detections by each survey. The gray, red, and blue histograms correspond to cuts with $q > 30$, 35, and 38 au, respectively.
Figure 12. Posterior absolute magnitude distributions of simulated detections. The red triangles represent the real ETNO detections by each survey. Note that ST has overlapping data points at $H = 6.7$. The gray, red, and blue histograms correspond to cuts with $q > 30$, 35, and 38 au, respectively.
Figure 13. Posterior longitude of ascending node distributions of simulated detections. The red triangles represent the real ETNO detections by each survey. The gray, red, and blue histograms correspond to cuts with \( q > 30, 35, \) and 38 au, respectively.
Figure 14. Posterior longitude of pericenter distributions of simulated detections. The red triangles represent the real ETNO detections by each survey. The gray, red, and blue histograms correspond to cuts with $q > 30$, 35, and 38 au, respectively.

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