Deep Gaussian Markov random fields

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Abstract

Gaussian Markov random fields (GMRFs) are probabilistic graphical models widely used in spatial statistics and related fields to model dependencies over spatial structures. We establish a formal connection between GMRFs and convolutional neural networks (CNNs). Common GMRFs are special cases of a generative model where the inverse mapping from data to latent variables is given by a 1-layer linear CNN. This connection allows us to generalize GMRFs to multi-layer CNN architectures, effectively increasing the order of the corresponding GMRF in a way which has favorable computational scaling. We describe how well-established tools, such as autodiff and variational inference, can be used for simple and efficient inference and learning of the deep GMRF. We demonstrate the flexibility of the proposed model and show that it outperforms the state-of-the-art on a dataset of satellite temperatures, in terms of prediction and predictive uncertainty.

1. Introduction

Convolutional neural networks (CNNs) are the de facto standard model in computer vision when training on a large set of images. Images are lattice-based data with local dependencies and thus have clear connections with spatial statistics. However, many spatial problems lack the abundance of data common to computer vision applications, and often we need to build a model based on a single “image”, i.e., data recorded over some spatial field. Models such as deep image prior (Ulyanov et al., 2018) have shown that CNN architectures can encode useful spatial priors even in such situations, but the dominant approach is still to model the spatial dependencies explicitly using, e.g., Gaussian processes (GPs) (Williams & Rasmussen, 2006) or Gaussian Markov random fields (GMRFs) (Rue & Held, 2005).

In this paper we show a formal connection between GMRFs applied to lattice data and CNNs. Common GMRFs based on nearest neighbour interactions can be viewed as special cases of a generative model where the inverse mapping from the spatial field \( x \) to a latent variable \( z \) is given by a 1-layer linear CNN. Since common GP models have previously been shown to be tightly linked with GMRFs (Lindgren et al., 2011), this connection applies to certain GPs as well.

Modeling the inverse mapping \( (x \rightarrow z) \) using a CNN results in an auto-regressive (AR) spatial model, whereas using a CNN for the forward mapping \( (z \rightarrow x) \) would correspond to a moving average (MA) model (see, e.g., (Ljung, 1999) for a discussion on AR and MA models in a time series context). This has the important implication that we obtain an infinite receptive field (i.e., infinite range on the spatial dependencies in \( x \)) even with a 1-layer model. Indeed, this is a well known property of GMRFs.

The interpretation of a GMRF as a 1-layer CNN opens up for a straightforward generalization to multi-layer architectures, resulting in deep GMRFs (DGMRFs). Even when all layers are linear this has important practical implications: adding layers corresponds to increasing the auto-regressive order of the GMRF, i.e., edges are added to the GMRF graph which improves its expressivity. For conventional GMRF algorithms, simply adding more edges can have have a significant impact on the computational complexity due to a reduced degree of sparsity of the resulting precision matrix. For a DGMRF, on the other hand, the structure imposed by a multi-layer CNN architecture results in a favorable computational scaling. Indeed, using variational inference for the latent variables, we obtain a learning algorithm that scales linearly both with the number of dimensions in the spatial field (“pixels”) and the number of layers (AR order). Furthermore, viewing GMRFs through the lens of deep learning allows us to use well-established toolboxes for, e.g., automatic differentiation and GPU training, to facilitate simple learning of DGMRFs.

After a review of GMRFs in Section 2, we introduce the DGMRF model in Section 3. Section 4 describes how to efficiently train the model, and how to compute the posterior predictive distribution, including uncertainty. We discuss related work in Section 5. The results in Section 6 illustrate how DGMRFs are adaptive models, with outstanding pre-
dictive ability. Section 7 concludes. Code for our methods and experiments are available at (insert github link).

2. Background

2.1. Gaussian Markov random fields

A Gaussian Markov random field (GMRF) \( x \) is an \( N \)-dimensional Gaussian vector with mean \( \mu \) and precision (inverse covariance) matrix \( Q \), so that \( x \sim \mathcal{N}(\mu, Q^{-1}) \). For each GMRF there exists a graph \( G = (\mathcal{V}, \mathcal{E}) \), with vertices \( \mathcal{V} \) that correspond to the elements in \( x \), and edges \( \mathcal{E} \) that define their conditional independencies. For example, the vertices could represent the pixels in an image, with the edges connecting neighboring pixels. Formally,

\[
\{i, j\} \notin \mathcal{E} \iff x_i \perp x_j \mid x_{-ij}, \text{ for all } i \neq j,
\]

where \( x_{-ij} \) refers to all elements except \( i \) and \( j \), meaning that two elements \( x_i \) and \( x_j \), that are not neighbors in the graph, are conditionally independent given the rest. Importantly, the edges \( \mathcal{E} \) also determine the zero-pattern in the precision matrix \( Q \), as every GMRF has the property

\[
\{i, j\} \in \mathcal{E} \iff Q_{ij} \neq 0, \text{ for all } i \neq j.
\]

This means that a sparsely connected graph \( G \) results in a sparse precision matrix \( Q \), which gives great computational gains compared to working with the dense covariance matrix in many large-scale applications.

2.2. Example: GMRF defined using convolution

As an example, consider the second-order intrinsic GMRF or thin-plate spline model (Rue & Held, 2005), which can be defined by \( x \sim \mathcal{N}(0, (G^\top G)^{-1}) \), with

\[
G_{ij} = \begin{cases} 
4 & \text{, for } i = j \\
-1 & \text{, for } i \sim j \\
0 & \text{, otherwise}
\end{cases}
\]

(1)

where \( i \sim j \) denotes adjacency\(^1\). Imputing missing pixel values conditioned on its second-order neighborhood using this prior is equivalent to bicubic interpolation. The non-zero elements of each row \( i \) of \( G \) and \( Q \), with respect to neighboring pixels in 2D, can be compactly represented through the stencils

\[
w_G : \begin{bmatrix} 
-1 & 4 & -1 \\
-1 & -8 & -1 \\
1 & 2 & 1
\end{bmatrix}, \quad w_Q : \begin{bmatrix} 
1 & -8 & 2 \\
2 & -8 & -2 \\
2 & -2 & 8
\end{bmatrix}.
\]

(2)

\(^1\)It is perhaps more standard to define \( G_{ij} \) equal to the number of neighbors of pixel \( i \), which makes a difference in the image border. Our definition is convenient here as it makes \( G \) invertible.

An equivalent definition of this model is to first define \( z \sim \mathcal{N}(0, I) \) and then \( x \) through the inverse transform

\[
z = G x.
\]

(3)

It is trivial that \( x \) is Gaussian with mean \( 0 \) and it can be readily verified that \( \text{Cov}(x) = G^{-1}I G^{-\top} = (G^\top G)^{-1} \).

In a third, equivalent definition the inverse transform \( z = G x \) is instead written as a convolution

\[
Z = \text{conv}(X, w_G),
\]

where \( Z \) and \( X \) are image representations of the vectors \( z \) and \( x \). The stencil \( w_G \) in Eq. (2) is here used as a filter and conv() denotes same convolution (padding=“SAME”), for the equivalence to hold. This observation, that a certain kind of GMRF can be defined with the convolution operator, a filter and an auxiliary variable is a key observation for DGMRFs, which are defined in Section 3.

2.3. Link between GMRFs and Gaussian processes

Another instructive view of GMRFs is as a representation of a Gaussian processes (GP) with Matrit kernel (Lindgren et al., 2011). This result comes from a stochastic partial differential equation (SPDE) of the form

\[
(\kappa^2 - \Delta)\gamma \tau u(s) = W(s),
\]

which can be shown to have a GP solution \( u(s) \) with Matrit covariance function (Whittle, 1954; 1963). Here, \( W(s) \) is Gaussian white noise in a continuous coordinate \( s \), \( \Delta \) is the Laplacian operator, and \( \kappa, \tau \) and \( \gamma \) are hyperparameters that appear in the Matrit kernel. In particular, \( \gamma \) controls the smoothness of the GP. Moreover, for positive integer values of \( \gamma \), a numerical finite difference approximation to the SPDE, on the integer lattice, is given by

\[
\tau ((\kappa^2 I + G)^\gamma x = z,
\]

(4)

with \( G \) defined as in Eq. (1). As in Eq. (3), this inverse transform describes a GMRF, here with precision matrix \( Q = \tau^2((\kappa^2 I + G)^\gamma)^\top((\kappa^2 I + G)^\gamma) \). Firstly, this provides a sparse representation of a GP as a GMRF, with a discrepancy that can be reduced by making the discretization of the SPDE finer. Secondly, it gives an interpretation of GMRFs as models with similar properties as GPs.

2.4. GMRFs as spatial priors

GMRFs are commonly used as spatial priors for latent variables, which is common in spatial statistics and image analysis. In the simplest case, the data \( y \) is modeled as Gaussian, and conditionally independent given \( x \)

\[
p(y|x) = \prod_{i=1}^M p(y_i|x_i), \quad y_i|x_i \sim \mathcal{N}(y_i|x_i, \sigma^2),
\]
where $M$ is the number of observed pixels. Typical problems include inpainting ($M < N$), and denoising ($\sigma^2 > 0$), where the target is to reconstruct the latent $x$. Conveniently, in this situation the GMRF prior is conjugate, so the posterior is also a GMRF

\[
x | y \sim \mathcal{N}(\tilde{\mu}, \tilde{Q}^{-1}), \quad \text{with} \quad \tilde{Q} = Q + \frac{1}{\sigma^2} I_m, \quad \tilde{\mu} = \tilde{Q}^{-1} \left( Q\mu + \frac{1}{\sigma^2} y_m \right).
\]

(5)

Here, the observation mask $m$ has value $0$ for missing pixels and $1$ elsewhere, $y_m$ are the observations with value $0$ at missing pixels and $I_m$ is the identity matrix with diagonal element $0$ at missing pixels. Although the posterior is on closed form, the computational cost associated with $Q^{-1}$, needed for the posterior mean and marginal standard deviations, can be high. This is discussed more in Section 4.

3. Deep Gaussian Markov random fields

3.1. Linear DGMRFs

We define a linear DGMRF using an auxiliary standard Gaussian vector $z$ and a bijective function $g_\theta : \mathbb{R}^N \rightarrow \mathbb{R}^N$,

\[
z \sim \mathcal{N}(0, I), \quad z = g_\theta(x).
\]

In other words, we define $x$ through the inverse transform $g_\theta^{-1}$. The function $g_\theta(x)$ is for now assumed to be linear, so we can write $g_\theta(x) = G_\theta x + b_\theta$, where $G_\theta$ is an invertible square matrix. The non-linear case is considered in Section 3.4. We will specify $g_\theta(x)$ using a CNN with $L$ layers. Let $Z_l$ be a tensor of dimension $H \times W \times C$, with height $H$, width $W$ and $C$ channels, and let $z_l = \text{vec}(Z_l)$ be its vectorized version, with length $N = HWC$. The output of layer $l$ is defined as

\[
Z_l = \text{conv}(Z_{l-1}, w_l) + b_l,
\]

(6)

where $w_l$ is a 4D-tensor containing $C \times C$ 2D-filters and $b_l$ is a set of $C$ biases. Here, $\text{conv}(\cdot)$ refers to multichannel same convolution, more details are given in Section 3.3. In particular, we define $Z_L \triangleq Z$ and $Z_0 \triangleq X$. The hyperparameters $\theta = \{ (w_l, b_l) : l = 1, \ldots, L \}$ will be omitted in the following for brevity.

Just as for normalizing flows (Dinh et al., 2014; Rezende & Mohamed, 2015), $g$ can be seen as a sequence of bijections $g = g_L \circ g_{L-1} \circ \cdots \circ g_1$, each with corresponding transform matrix $G_l$. Since $g$ is linear, $x$ is a GMRF with density

\[
p(x) = \frac{\lvert \det(G) \rvert}{(2\pi)^{N/2}} \exp \left( -\frac{1}{2} (x - \mu)^	op G^	op G (x - \mu) \right),
\]

(7)

with $G = G_L G_{L-1} \cdots G_1$ and the mean $\mu = -G^{-1}b$ where $b$ can be computed as $b = g(0)$. This GMRF has precision matrix $Q = G^	op G$, that is guaranteed to be positive (semi-)definite for all $\theta$, which gives an advantage compared to modeling $Q$ directly.

The reason for defining $x$ through an inverse transform $z = g(x)$, rather than a forward transform, is twofold. Firstly, it establishes a formal connection to traditional GMRFs, see Proposition 1 below. Secondly, it gives an AR model with infinite receptive field, meaning that the output prediction at each pixel depends on all the pixels, even for a one-layer model. Compared to dilated CNNs (e.g. Yu & Koltun, 2015), which achieve long (yet finite) range dependencies through several layers of dilated filters, this is a simpler construction.

3.2. Computationally cheap convolutional filters

In this paper, we consider two special forms of convolutional filters: plus ($+$) filters and sequential (seq) filters, with forms

\[
+ : \begin{bmatrix} a_3 \\ a_2 \\ a_1 \end{bmatrix} \quad \text{seq} : \begin{bmatrix} a_3 \\ a_2 \\ a_1 \end{bmatrix}
\]

(8)

where $a_1, \ldots, a_5 \in \mathbb{R}$ are parameters to be learned, and the empty positions are fixed to zero. The benefit of filters with these special structures is that they correspond to linear transforms for which the determinants $\det(G_l)$ in

\[
\det(G) = \prod_{l=1}^L \det(G_l)
\]

(9)

needed in Eq. (7), can be cheaply computed.

Defining a linear DGMRF through a sequence of convolutions using different small filters is of course equivalent to defining it using a single convolution with a larger filter, apart from at the boundary. The equivalent larger filter can be obtained by sequentially convolving the smaller filters with themselves. The main motivation for using the deep definition is that it has cheap determinant calculations, using Eq. (9) and the $+$- and seq-filters, which is not the case for a general larger filter. Also, a deep definition results in fewer parameters to learn. For example, four full $3 \times 3$-filters has 36 parameters compared to 81 parameters for the corresponding $9 \times 9$-filter. Finally, the deep convolutional architecture has proven successful in countless applications of neural networks, which makes it plausible that also GMRFs should benefit from the same inductive biases. The addition of non-linearities between the layers is discussed in Section 3.4. We now discuss the two filter types and their determinant computation for the singlechannel case and discuss the multichannel construction in Section 3.3.

3.2.1. $+$-FILTERS

We begin with two propositions, the first connecting linear DGMRFs with $+$-filters to the GMRFs in Section 2, and the second giving the cheap determinant computation.

Proposition 1. The second-order intrinsic GMRF, as well
as the GMRF approximation of a Matrn GP, are special cases of the linear DGMRF model with $+\text{-filters}$.

Proof. The non-zero pattern of the $+\text{-filter}$ is the same as that of $w_{\phi}$ in Eq. (2), meaning that a one-layer linear DGMRF with the same filter weights is equivalent to the second-order intrinsic GMRF. Similarly, Eq. (4) can be written as a linear DGMRF with $L$ layers of $+\text{-filters}$, for $L \geq \gamma$. This requires that $\gamma$ of the layers have the same filter with $a_1 = 4 + \sqrt{2}$ and $a_2 = \cdots = a_5 = -1$, and the other $L - \gamma$ layers to be identity functions.

Proposition 2. The linear transform matrix $G_+$ defined by some convolution of an $H \times W$ image with the $+\text{-filter}$ defined in Eq. (8), has determinant

$$\det (G_+) = \prod_{i=1}^{H} \prod_{j=1}^{W} \left[ a_1 + 2 \sqrt{a_3 a_5} \cos \left( \pi \frac{i}{H+1} \right) + 2 \sqrt{a_2 a_4} \cos \left( \pi \frac{j}{W+1} \right) \right],$$

where $\sqrt{-1}$ is treated as imaginary. Computing the determinant thus has complexity $O(N)$.

The proof, given in detail in the supplement, is to show that the factors of this product are identical to the eigenvalues of $G_+$, which can be done by writing $G_+$ as a Kronecker sum of tridiagonal Toeplitz matrices. Proposition 2 provides a fast method for computing the determinant, which would have complexity $O(N^3)$ for a general square matrix, or $O(N^{3/2})$ if based on sparse Cholesky factorization (Rue & Held, 2005). In practice, we reparameterize $a_1, \ldots, a_5$ to ensure that all the eigenvalues are real and positive, see details in the supplement. Without this constraint, we have observed unstable, oscillating solutions in some situations, and the constraint also ensures that the bijective assumption is valid.

3.2.2. seq-filters

For some ordering of the image pixels, the output pixels for a convolution with a seq-filter only depend on previous input pixels. This is equivalent to saying that the corresponding transform matrix $G_{\text{seq}}$ yields a lower triangular matrix $P^T G_{\text{seq}} P$ for some permutation matrix $P$, which implies that $\det(G_{\text{seq}}) = a_1^N$. Seq-filters are therefore extremely cheap, but somewhat restricted compared to $+\text{-filters}$, due to the inability to model non-sequential patterns. However, the seq-filter in Eq. (8) can be rotated/mirrored in eight different ways, each encoding a different pixel ordering. Thus different sequential dependencies can be modelled in different layers. Also seq-filters can trivially be extended to $5 \times 5$- or $7 \times 7$-filters, still without changing the determinant. Connections to auto-regressive models, for example PixelCNN (Van den Oord et al., 2016b), are discussed in Section 5.

3.3. Multichannel construction

When $C > 1$, Eq. (6) can be written on vector form as

$$z_l = G_l z_{l-1} + b_l,$$

where $G_l$ is the transition matrix of a single convolution from input channel $j$ to output channel $i$, and $b_{l,j}$ is the bias of output channel $j$. In order to make $\det(G_l)$ computationally tractable, we set $G_{l,i,j} = 0$ for $i < j$, making $G_l$ lower block triangular and

$$\det (G_l) = \prod_{c=1}^{C} \det (G_{l,c,c}).$$

The ordering of channels could vary between different layers to allow information to flow back and forth between all channels. One could also add invertible $1 \times 1$ convolutions (Kingma & Dhariwal, 2018) between layers for this ordering to be dynamic and learnable. A multichannel DGMRF could learn more interesting representations by storing different features in different channels of the hidden layers. Even with singlechannel data, the hidden layers can be multichannel by using a multiscale architecture (Dinh et al., 2017).

3.4. Non-linear extension

The linear DGMRFs can be extended by adding non-linear activation functions between layers of the neural network $g_{\theta}(x)$. Formally, Eq. (6) is replaced by

$$Z_l = \psi_l \left( \text{conv}(Z_{l-1}, W_l) + b_l \right),$$

where $\psi_l$ is a non-linear scalar function that operates element-wise. We restrict $\psi_l$ to be strictly increasing, to ensure that $g_{\theta}(x)$ is a bijection.

The distribution of $x$ can now be computed by the change of variable rule

$$\log p(x) = \log p(z) + \log |\det(dz/dx)|$$

$$= \log p(z) + \sum_{l=1}^{L} \log |\det(G_l)| + \sum_{l=1}^{L} \sum_{i=1}^{N} \log |\psi_l'(h_{l,i})|,$$

where $h_l = \text{vec}(\text{conv}(Z_{l-1}, W_l) + b_l)$ is the input to $\psi_l$ and $\psi_l'$ is the derivative. Just as in the linear case, the computational cost of this density is linear.

Per default we assume that $\psi_l$ are Parametric Rectified Linear Units (PReLUs), defined by

$$\psi_l(h) = \left\{ \begin{array}{ll} h, & \text{for } h \geq 0 \\ \alpha_l h, & \text{for } h < 0 \end{array} \right.,$$

where $\alpha_l$ are learnable parameters with $\alpha_l > 0$. We can now add $\alpha_1, \ldots, \alpha_L$ to the hyperparameters $\theta$ and optimize them jointly.
4. Learning and inference

There exist two kinds of parameters that we wish to infer: the latent field $x$ and the hyperparameters $\theta$. Since we are not interested in the posterior uncertainty of $\theta$ directly, we take a practical course of action and optimize these, using a variational lower bound on the marginal likelihood $p(y|\theta)$. Given the optimal value $\hat{\theta}$, for the linear model, we make a fully Bayesian treatment of the posterior $p(x|\hat{\theta}, y)$.

4.1. Optimization of hyperparameters

For the linear DGMRF, the marginal likelihood can be written on closed form as

$$p(y|\theta) = \frac{p(y|x, \theta) p(x|\theta)}{p(x|y, \theta)}|_{x=x^*}$$

for arbitrary value of $x^*$. Unfortunately, this expression requires the determinant of the posterior precision matrix $\det(G^T G + \sigma^{-2} I_m)$, which is computationally infeasible for large $N$. Instead, we focus on the variational evidence lower bound (ELBO) $\mathcal{L}(\theta, \phi, y) \leq \log p(y|\theta)$ which can be written as

$$\mathcal{L}(\theta, \phi, y) = \mathbb{E}_{q_\phi(x)} \left[ - \log q_\phi(x) + \log p(y, x|\theta) \right],$$

where $q_\phi(x)$ is the variational posterior approximation, which depends on variational parameters $\phi$. We here only intend to use $q_\phi(x)$ as a means for optimizing $\theta$, and not for example to make posterior predictions. For simplicity, we choose $q_\phi(x) = \mathcal{N}(x|\nu, S)$ with diagonal covariance matrix. After inserting the variational and model densities and simplifying, the ELBO can be written as

$$\mathcal{L}(\theta, \phi, y) = \frac{1}{2} \log | \det (S_\phi) | - M \log \sigma_\theta + \log | \det (G_\theta) |$$

$$- \frac{1}{2} \mathbb{E}_{q_\phi(x)} \left[ g_\phi(x)^T g_\phi(x) + \frac{1}{\sigma^2_\theta} (y_m - x)^T I_m (y_m - x) \right],$$

where constant terms have been omitted, and all parameters have been subscripted with $\theta$ and $\phi$ to clarify whether they are model or variational parameters. Additionally, we use the reparameterization trick (Kingma & Welling, 2013) replacing the last expectation with a sum over $N_\eta$ standard random samples $\varepsilon_1, \ldots, \varepsilon_{N_\eta}$, and set $x_i = \nu_\phi + S_\phi^{1/2} \varepsilon_i$ in the sum. This gives an unbiased estimator of the ELBO, which has low variance. Moreover, this estimator is differentiable with respect to $\theta$ and $\phi$ and can be used for stochastic gradient optimization with autodiff and backprop. Hyperparameter learning will be fast, with a time complexity that is $O(N)$ for a fixed number of iterations of optimization, since backprop in a CNN is linear and so are the determinant computations described in Section 3.2. This can be compared with $O(N^3)$ for standard GP and $O(N^{3/2})$ for standard GMRF inference in 2D problems, based on the Cholesky-decomposition. The ELBO can be trivially extended to the non-linear model by replacing $\log | \det (G_\theta) |$ in Eq. (10).

4.2. Exact inference for the latent field

For a linear DGMRF, the conditional posterior $p(x|\hat{\theta}, y)$ is also a GMRF, see Eq. (5). Even though computing this density is too costly in general, it is possible to compute the posterior mean and to draw samples from the posterior, which can be used for making predictions. Both require solving linear equation systems $Qx = c$ involving the posterior precision matrix $Q = G^T G + \sigma^{-2} I_m$. For this we use the conjugate gradient (CG) method (see e.g. Barrett et al., 1994), which is an iterative method that, rather than exactly computing $x = Q^{-1} c$, iteratively minimizes the relative residual $||Qx - c||/||c||$ until it falls below some threshold $\delta$, that can be set arbitrarily low. In practice, $\delta = 10^{-7}$ gives solutions that visually appear to be identical to the exact solution. CG only requires matrix-vector-multiplications, which means that the multiplications with $G$ and $G^T$ can be implemented in a matrix-free fashion using convolution.

Posterior sampling for $x$ can be performed using the method of Papandreou & Yuille (2010) by computing

$$x_s = \tilde{Q}^{-1} \left( G^T (u_1 - b) + \frac{1}{\sigma^2} (y_m + \sigma u_2 \odot m) \right),$$

where $u_1$ and $u_2$ are standard Gaussian random vectors, and $\odot$ denotes element-wise product. It can be easily verified that $x_s$ is Gaussian, with the same mean and covariance as the posterior.

Given a number of posterior samples, the posterior marginal variances $\text{Var}(x_i|y, \hat{\theta})$ can be naively approximated using Monte Carlo estimation, but more efficiently approximated using the simple Rao-Blackwellized Monte Carlo (simple RBMC) method by Sidén et al. (2018). This provides a way to compute the posterior predictive uncertainty as

$$\text{Var}(y_i^*|y) = \text{Var}(x_i|y, \hat{\theta}) + \sigma^2.$$
tional approximation \( q_\phi(x) \) instead, but the proposed Gaussian independent variational family is probably too simple to approximate the true posterior in a satisfying way. We discuss possible solutions to this limitation in Section 7.

5. Related work

5.1. GMRFs

GMRFs have a long history in spatial and image analysis (see e.g. Woods, 1972; Besag, 1974 and Rue & Held, 2005). The mentioned SPDE approach has been extended in numerous ways leading to flexible GMRF models, with for example non-stationarity (Fuglstad et al., 2015), oscillation (Lindgren et al., 2011) and non-Gaussianity (Bolin, 2014).

A nested (deep) SPDE is introduced by Bolin & Lindgren (Wilson & Nickisch, 2015) unless P (Dunlop et al., 2018). However, standard GPs are limited by Stein (1999; Williams & Rasmussen, 2006) are commonplace for modelling spatially dependent data, not necessarily restricted to observations on the grid. The GP covariance kernel, which encodes the spatial dependencies, can be made flexible (Wilson & Adams, 2013) and deep (Dunlop et al., 2018). However, standard GPs are limited by \( O(N^3) \) computational complexity (assuming the number of measurements \( M \sim O(N) \)).

Inducing point methods (Quiñonero-Candela & Rasmussen, 2005; Snelson & Ghahramani, 2006) can reduce the complexity to \( O(P^2 N + P^3) \) or even \( O(P^3) \) (Hensman et al., 2013), where \( P \) is the number of inducing points. However, for grid data, these methods tend to over-smooth the data (Wilson & Nickisch, 2015) unless \( P \) is chosen in the same order of magnitude as \( N \). When the data are on a regular grid and fully observed, Kronecker and Toeplitz methods can be used for fast computation (e.g. Saatçi, 2011; Wilson, 2014). Under certain assumptions about the interactions across the input dimensions, such as additivity and separability, this can reduce the complexity to \( O(N \log N) \). These methods can also be extended to when the data are not on the grid (Wilson & Nickisch, 2015), or when the grid is incomplete (Stroud et al., 2016).

5.3. Deep generative models

Deep generative models are normally trained on large datasets, and can then be used to generate new samples from the data distribution, or to explore learned latent representations. For example, generative adversarial networks (GANs) have been used for inpainting problems in natural images (see e.g. Yu et al., 2018; Liu et al., 2018) with impressive results.

Invertible generative models, e.g., normalizing flows (Rezende & Mohamed, 2015), real NVP (Dinh et al., 2017), Glow (Kingma & Dhariwal, 2018) and i-ResNet (Behrmann et al., 2019), model \( x \) as a bijective function of a latent variable, just as our model. The likelihood can be optimized directly for hyperparameter learning, but generally require all pixels to be non-missing. Dinh et al. (2014) use a learned model for inpainting with projected gradient ascent, which gives the posterior mode but no uncertainty quantification.

Auto-regressive models, e.g. PixelRNN (Van den Oord et al., 2016a), PixelCNN (Van den Oord et al., 2016b) and PixelCNN++ (Salimans et al., 2017), model the pixel values sequentially similarly to our proposed seq-filters, but there are some major differences. The seq-filters use masked convolutions to obtain cheap determinant computations and the ordering of pixels can change between layers, whereas auto-regressive models have no latent variables and are trained using the same image as input and output. (Van den Oord et al., 2016a) consider image completion, but are limited to the e.g. the case where the bottom half of the image is missing, and can’t handle a general mask.

Variational autoencoders (Kingma & Welling, 2013) differ from our model in that they use the deep transform of the latent variables for modeling the parameters of the distribution of the data, rather than for the data directly. This makes the recovery of the latent variables intractable, which is not the case for our model, however, we still use the same variational optimization algorithm for speedups.

Deep image prior (Ulyanov et al., 2018) uses CNNs, trained on single images, for denoising and inpainting, with promising results. However, this is not truly a probabilistic generative model, since the latent variables are not inferred, but fixed to some initial random values. Thus, it would be difficult to output predictive uncertainty from such a model.
Figure 1. Posterior mean for inpainting the 160×120 pixel toy data without edges (top) and with edges (bottom). The second column shows inpainting by the same Matrn GMRF model and hyperparameters that were used to generate the data without edges.

6. Experiments

We have implemented DGMRF in TensorFlow (Abadi et al., 2016), taking advantage of autodiff and GPU computations. We train the hyperparameters, and compute the predictive distribution using the CG algorithm. To avoid boundary effects, the images are extended with a 10-pixel wide frame of missing values at all sides. Additional details about the implementation can be found in the supplement.

6.1. Toy data

We demonstrate the behaviour of our model for an inpainting problem on the two toy datasets in Figure 1, which have size 160×120 pixels. The data in the first row are generated as a random sample from the Matrn GMRF in Eq. (4) with \( \gamma = 1, \tau = 1, \) and \( \kappa^2 = 8 \times 50^2 \) corresponding to a spatial correlation range of 50 pixels. The data in the second row are the same, but with added horizontal and vertical edges. Column 2 shows the posterior mean for \( x \) when the Matrn model with the same parameters is used for prediction, which gives optimal results for the first row, but which over-smoothes the edges in the second row. The corresponding results for different instances of the linear DGMRF, are shown in column 3-6. They all perform well in the simple case without edges, in which the +−-filter models contain the true model according to Proposition 1. In the case with edges, the model with depth \( L = 1 \) which corresponds to a standard GMRF is too simple to handle the more complex structure in the data, but for \( L = 3 \), all the models give reasonable results, that preserve the edges.

Figure 2 displays the learned +−-filters of the 3-layer model, and the values of valid pixels of the hidden layers, when original data, including missing pixels, are used as input. The first two filters learn differentiation in the vertical and horizontal direction, and the third filter is close to an identity function. Most spatial structure is removed in layer 3, that is assumed to be standard normal by the model.

6.2. Satellite data

We compare our method against some popular methods for large data sets in spatial statistics, by considering the satellite data of daytime land surface temperatures, used in the competition by Heaton et al. (2018). The data are on a 500×300 grid, with 105,569 non-missing observations as...
We have proposed deep GMRFs which enable us to view where it is slightly worse than NNGP. In terms of MAE, where 

\[
y_i \mid x_i \sim N \left( y_i \mid x_i + F_i, \beta, \sigma^2 \right),
\]

where \( F \) is a spatial covariate matrix with columns corresponding to (constant, longitude, latitude), and \( \beta \) is a 3-dimensional regression coefficient vector, which can be integrated out jointly with \( x \) for the predictions, see details in the supplement.

Table 1 compares different instances of our model with the methods in the competition. The scores used are mean absolute error (MAE), root-mean-squared-error (RMSE), mean continuous rank probability score (CRPS), mean interval score (INT), and prediction interval coverage (CVG). CRPS and INT are proper scoring rules, that also account for the predictive uncertainty (Gneiting & Raftery, 2007). For the DGMRFs, based on the first three scores, the seq-filters of size 5 × 5 perform better compared to those of size 3 × 3 and the + - filters, which may be due to the increased flexibility of larger filters. Moreover, deeper models tend to give better results. For the non-linear (NL) model, CG cannot be used to compute posterior mean and uncertainty, and the first two scores are instead computed based on the mean of the variational approximation. We note that the NL model performs worse than the linear. DGMRF outperforms all the methods from the competition on all criteria except CVG, which also account for the predictive uncertainty (Gneiting & Raftery, 2007). For the DGMRFs, based on the first three scores, the seq-filters of size 5 × 5 perform better compared to those of size 3 × 3 and the + - filters, which may be due to the increased flexibility of larger filters. Moreover, deeper models tend to give better results. For the non-linear (NL) model, CG cannot be used to compute posterior mean and uncertainty, and the first two scores are instead computed based on the mean of the variational approximation. We note that the NL model performs worse than the linear. DGMRF outperforms all the methods from the competition on all criteria except CVG, which also account for the predictive uncertainty (Gneiting & Raftery, 2007).

Table 1. Prediction scores on the satellite data. The scores of the methods in the upper pane come from Table 3 in Heaton et al. (2018). Our models are presented in the lower pane. Lower scores are better, except for CVG, for which 0.95 is optimal. The results for our models are averages computed across five random seeds. Standard deviations across seeds are shown in parenthesis for seq_{5 × 5, L=5} and for the other models in the supplement.

| Method      | MAE  | RMSE | CRPS  | INT  | CVG  |
|-------------|------|------|-------|------|------|
| FRK         | 1.96 | 2.44 | 1.44  | 14.08| 0.79 |
| Gapfill     | 1.33 | 1.86 | 1.17  | 34.78| 0.36 |
| LatticeKrig | 1.22 | 1.68 | 0.87  | 7.55 | 0.96 |
| LAGP        | 1.65 | 2.08 | 1.17  | 10.81| 0.83 |
| MetaKriging | 2.08 | 2.50 | 1.44  | 10.77| 0.89 |
| MRA         | 1.33 | 1.85 | 0.94  | 8.00 | 0.92 |
| NNGP        | 1.21 | 1.64 | 0.85  | 7.57 | 0.95 |
| Partition   | 1.41 | 1.80 | 1.02  | 10.49| 0.86 |
| Pred. Proc. | 2.15 | 2.64 | 1.55  | 15.51| 0.83 |
| SPDE        | 1.10 | 1.53 | 0.83  | 8.85 | 0.97 |
| Tapering    | 1.87 | 2.45 | 1.32  | 10.31| 0.93 |
| Peri. Embe. | 1.29 | 1.79 | 0.91  | 7.44 | 0.93 |
| Deep GMRF   |      |      |       |      |      |
| seq_{5 × 5, L=1} | 1.06 | 1.42 | 0.76  | 7.21 | 0.97 |
| seq_{5 × 5, L=3} | 0.95 | 1.3  | 0.75  | 8.29 | 0.97 |
| seq_{5 × 5, L=5} | **0.93** | **1.25** | **0.74** | **8.14** | 0.97 |
| seq_{1 × 3, L=5} | **1.16** | **1.57** | **0.81** | **6.98** | 0.97 |
| seq_{1 × 5, L=5} | **1.09** | **1.47** | **0.78** | **7.63** | 0.97 |
| seq_{5 × 5, L=5, NL} | **1.37** | **1.87** | -    | -   | -   |

7. Conclusions and future work

We have proposed deep GMRFs which enable us to view (high-order) GMRF models as CNNs. We have focused our attention on lattice-based graphs to clearly show the connection to conventional CNNs, however, the proposed method can likely be generalized to arbitrary graphs via graph convolutions (see, e.g., Xu et al., 2019), and to continuously referenced data, similar to Bolin & Lindgren (2011). We have also primarily considered linear CNNs, resulting in a model that is indeed equivalent to a GMRF. The DGMRF nevertheless has favorable properties compared to conventional algorithms for GMRFs: the CNN architecture opens up for simple and efficient learning, even thought the corresponding graph will have a high connectivity when using multiple layers of the CNN. Empirically we have found that multi-layer architectures are indeed very useful even in the linear case, enabling the model to capture complex data dependencies such as distinct edges, and result in state-of-the-art performance on a spatial benchmark problem.

Using a CNN for the inverse mapping when defining the DGMRF results in a spatial AR model. This gives rise to an infinite receptive field, but can also result in instability of the learned prior. We have constrained the filter parameters to yield real eigenvalues, and under this constraint we have not seen any issues with instability. A more principled way of addressing this potential issue is left for future work.

The CNN-based interpretation offers a natural extension, namely to introduce non-linearities to enable modeling of more complex distributions. Further exploring this extension requires future work. In particular we believe that the independent variational approximation is insufficient in non-linear DGMRFs. One way to address this limitation is to parameterize the square-root of the covariance matrix \( S_{\phi}^{1/2} \) of the variational approximation \( q_{\phi} \) directly as a lower triangular matrix. Another approach is to model \( q_{\phi} \) as a GMRF with the same graph structure as the original model.
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Supplementary material

Proof of Proposition 2

The transform matrix $G_{+}$ can be written as

$$G_{+} = T_1 \oplus T_2 = T_1 \otimes I_H + I_W \otimes T_2,$$

where $\oplus$ denotes the Kronecker sum and $\otimes$ the Kronecker product, $T_1$ is a $W \times W$ tridiagonal Toeplitz matrix, denoted $T_1 = (W; a_2, a_1/2, a_4)$, meaning that

$$T_1 = \begin{bmatrix}
    a_1/2 & a_4 & 0 \\
    a_2 & a_1/2 & a_4 \\
    0 & a_2 & a_4 \\
    \vdots & \ddots & \ddots \\
    \vdots & & \ddots & a_4 \\
    0 & a_2 & a_4 & a_1/2
\end{bmatrix},$$

and similarly $T_2 = (H; a_3, a_1/2, a_5)$. For the eigenvalues of a Kronecker sum, it holds that if $\lambda_1$ is an eigenvalue of $T_1$ and $\lambda_2$ is an eigenvalue of $T_2$, then $\lambda_1 + \lambda_2$ is an eigenvalue of $T_1 \oplus T_2$ (Graham, 1981). Moreover, the eigenvalues of a tridiagonal Toeplitz matrix $T = (n; b, a, c)$ have a simple formula

$$\lambda_i (T) = a + 2\sqrt{bc} \cos \left( \frac{\pi i}{n+1} \right), \quad \text{for } i = 1, \ldots, n,$$

which holds for real and complex $a$, $b$ and $c$ (Smith, 1985). Substituting this formula into the expression

$$\det(G_{+}) = \prod_{i=1}^{H} \prod_{j=1}^{W} (\lambda_i (T_2) + \lambda_j (T_1))$$

gives the result in Proposition 2.

$+$-filter reparameterization

The following reparameterization is used to ensure that $G_{+}$ has real positive eigenvalues

$$a_1 = \text{softplus} (\rho_1) + \text{softplus} (\rho_2)$$

$$a_2 a_4 = (\text{softplus} (\rho_1) \tanh (\rho_3) / 2)^2, \quad a_4/a_2 = \exp (\rho_4)$$

$$a_3 a_5 = (\text{softplus} (\rho_2) \tanh (\rho_5) / 2)^2, \quad a_5/a_3 = \exp (\rho_6),$$

where $\rho_1, \ldots, \rho_6$ are real numbers.

Implementation details

The hyperparameters $\theta$ and variational parameters $\phi$ are trained with respect to the negative ELBO (Eq. (10)) divided by $N$ as loss function, using Adam optimization (Kingma & Ba, 2014) with default settings, learning rate 0.01 and 100k iterations. The parameters with the lowest loss value are then saved and conditioned on by our implementation of the CG algorithm, for computing the posterior mean and standard deviation of $x$. We use $N_q = 10$ samples from variational approximation to compute the expectation in each iteration. We can train the measurement error $\sigma$ together with the other parameters $\theta$, but we have used a fixed $\sigma = 0.001$, which seems to give very similar results, but with somewhat faster convergence. For the DGMRFs with seq-filters, we randomly select among the eight possible orientations of the filters in each layer. As the toy data is centered around 0, the bias in each layer was fixed to 0 for this experiment. The satellite data was normalized to have maximum pixel value 1.

Linear trend model

For inference with the linear trend model in Eq. (11), we extend the vector of latents $x$ to include also the regression coefficients $\beta$, and use (for linear DGMRFs) the prior

$$\begin{bmatrix} z \\ z' \end{bmatrix} = \begin{bmatrix} G & 0 \\ 0 & vI \end{bmatrix} \begin{bmatrix} x \\ \beta \end{bmatrix} \leftrightarrow \bar{z} = G \bar{x}, \quad \bar{z} \sim \mathcal{N}(0, I),$$

where $v$ can be interpreted as the prior inverse standard deviation of the elements of $\beta$, which we fix at $v = 0.0001$. The posterior for $\bar{x}$ is a GMRF, similar to Eq. (5), with

$$\bar{Q} = G^\top G + \frac{1}{\sigma^2} \begin{bmatrix} I_F^\top \\ F^\top \end{bmatrix} \begin{bmatrix} I_m & \begin{pmatrix} I & F \end{pmatrix} \end{bmatrix},$$

$$\bar{\mu} = \bar{Q}^{-1} \begin{pmatrix} b \\ \begin{pmatrix} 0 \\ \begin{pmatrix} 1 & F^\top \end{pmatrix} y_m \end{pmatrix} \end{pmatrix},$$

and thus we can proceed with inference as before, with $\bar{x}$ instead of $x$, with slight modifications to the ELBO and to the CG method. We use an independent variational approximation $q_{\phi \theta}(\beta) = \mathcal{N}(\beta|\nu_\beta, S_\beta)$ for $\beta$. Integrating out $\beta$ is important for the predictive performance. For reference, if linear trends are instead removed using the ordinary least squares estimates of $\beta$ in a preprocessing step, the row in Table 1 corresponding to seq5x5,L=5 instead reads (1.25, 1.74, 0.90, 8.45, 0.89). When the linear trend model is used, we compute posterior standard deviations using standard Monte Carlo estimates, instead of simple RBMC, using $N_s = 100$ samples.
Figure 3. Satellite data inpainting by a linear DGMRF with 5 layers of 5×5 seq-filters.

| Method          | MAE   | RMSE  | CRPS  | INT   | CVG   |
|-----------------|-------|-------|-------|-------|-------|
| seq5×5, L=1     | 0.029 | 0.040 | 0.011 | 0.216 | 0.000 |
| seq5×5, L=3     | 0.022 | 0.042 | 0.019 | 0.462 | 0.001 |
| seq5×5, L=5     | 0.037 | 0.051 | 0.012 | 0.461 | 0.001 |
| seq3×3, L=5     | 0.066 | 0.097 | 0.039 | 0.171 | 0.003 |
| seq5×5, L=5, NL | 0.039 | 0.056 | 0.018 | 0.221 | 0.001 |
| data            | -     | -     | -     | -     | -     |
Supplement references

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