Survival comparison between radiofrequency ablation and surgical resection for patients with small hepatocellular carcinoma: A systematic review and meta-analysis

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Abstract

Objective: This study aimed to evaluate and compare the long-term therapeutic efficacy of radiofrequency ablation (RFA) versus that of surgical resection in small hepatocellular carcinoma (HCC).

Methods: Relevant articles in English from PubMed, EMBASE, and the Cochrane Library were retrieved. Pooled hazard ratios (HRs) were calculated to assess the prognostic value of RFA compared with that of surgical resection.

Results: A total of 19 studies involving 15,071 patients were included. The combined HRs (95% confidence interval [CI]) of RFA for recurrence/relapse-free survival (RFS) and overall survival (OS) were 1.55 (95% CI = 1.29–1.86, I² = 72.5%) and 1.61 (95% CI = 1.29–2.01, I² = 60.4%), respectively, compared with surgical resection. In subgroup analyses according to study design, both RFS and OS of the prospective subgroups showed statistical significance, and no statistical heterogeneity existed between studies.

Conclusion: Our clinical data suggest that surgical resection offers better long-term oncologic outcomes than RFA.

Abbreviations: HCC = hepatocellular carcinoma, HR = hazard ratio, OS = overall survival, RFA = radio frequency ablation, RFS = recurrence/relapse-free survival.

Keywords: hepatocellular carcinoma, long-term efficacy, meta-analysis, radiofrequency ablation, surgical resection

1. Introduction

Liver cancer is the third leading cause of mortality worldwide and includes hepatocellular carcinoma (HCC) and other primary liver cancers (for example, intrahepatic cholangiocarcinoma).[1]

Treatment options for HCC currently include liver transplantation, surgical resection, radiofrequency ablation (RFA), transarterial chemoembolization, and drug therapy.[2] Based on the current guidelines of the European Association for the Study of the Liver and American Association for the Study of Liver Diseases, surgical resection and RFA, along with liver transplantation, are the recommended treatment modalities for very small HCC, based on the presence or absence of portal hypertension and associated diseases.[3,4]

While liver transplantation is regarded as an ideal choice for patients with small HCC, this option is significantly limited due to organ shortages. According to the Barcelona Clinic Liver Cancer (BCLC) system, early HCC (BCLC 0 or A) should be treated with surgical resection or RFA.[5] Although surgical resection causes greater damage to the non-tumor liver parenchyma, it is generally preferred for small HCCs because surgical resection offers the therapeutic possibility of complete eradication of satellite tumor lesions and microscopic tumor emboli in the adjacent vasculature.[6–9] RFA, which involves heating the tumor and surrounding liver tissue, is generally accepted as a priority treatment in patients with impaired liver functional reserve because of its excellent efficacy and safety as well as better tolerability.[7–9]

To date, several studies have compared the curative effects of RFA and surgical resection for HCC. The optimal management choice considering long-term overall and recurrence/relapse-free survival for patients with small HCC is a matter of debate. Several researchers have reported higher overall or disease-free survival in patients who underwent surgical resection compared to that in patients who underwent RFA.[10–12] In contrast, no such correlation was observed by Lee et al (2017).[13] To clarify
this issue, a meta-analysis was designed to compare the long-term oncologic outcomes of RFA and surgical resection in patients with small HCC.

2. Materials and methods

2.1. Registration

This systematic review including a meta-analysis was prospectively registered with the PROSPERO International Prospective Register of Systematic Reviews (PROSPERO identifier CRD42020152746). All analyses were based on previously published studies; thus, no ethical approval and patient consent are required.

2.2. Inclusion criteria and literature source retrieval strategy

A systematic search of PubMed, Embase, and the Cochrane Library (2006–2017) was performed. The following search items were listed: (“hepatocellular carcinoma” [All Fields]) AND (“percutaneous local ablative therapy” [All Fields] OR “radiofrequency ablation” [All Fields]) AND (“hepatectomy” [All Fields] OR “resection” [All Fields]). To be considered eligible for this meta-analysis, the enrolled studies were required to include comparative data on clinical efficacy (overall survival [OS] and adverse events of HCC [RFS]). The inclusion criteria were as follows:

(1) trial design: studies with comparative data on clinical effectiveness and adverse events of RFA and surgical resection for HCC;
(2) clearly documented indications for use of RFA and surgical resection;
(3) treatment design: RFA vs surgical resection or surgical resection vs. RFA; and
(4) characteristics of patients.

Studies were required to include relatively integrated information (demography and basic characteristics) on the enrolled patients, such as average age, percentage of males, Child-Pugh class, and tumor size. Abstracts, letters, editorials and expert opinions, reviews without original data, case reports, and studies lacking control groups were excluded in addition to studies that focused on unresectable HCC or HCC recurrence after hepatectomy, had no clearly reported outcomes of interest, evaluated patients with cholangiocellular carcinomas or liver metastases, or employed a combination of surgical resection and ablation in patients with HCC.

2.3. Data extraction

Data extraction was independently conducted by two reviewers using standardized methods. Any disagreements were settled by discussion of the relevant study data and adjudicated by an experienced reviewer (Table 1). From each study, the following data were extracted: publication details (name of the first author and year of publication) and study characteristics (study design, age, gender, study design, tumor size, Child-Pugh class, endpoints).

2.4. Statistical analysis

In this meta-analysis, recurrence/relapse-free survival (RFS), disease-free survival (DFS), progression-free survival (PFS), and tumor-free survival (TFS) in the included studies were combined and redefined as RFS. DFS was defined as the period after successful treatment in which there were no symptoms or effects of the disease. For PFS analysis, follow-up durations were estimated from the date of initial treatment to the date of disease progression. TFS was defined as the time from the start of treatment to the appearance of recurrence or metastasis. In general, all these terms are not synonymous but can at times represent the same outcomes. These outcomes can be reflected as local or systemic recurrences. Overall survival (OS) was defined as the time from therapy initiation until death regardless of the cause. In terms of the effect size of each study, hazard ratio (HR) and 95% confidence interval (CI), taking into account the number and time of events, were calculated. HR referred to the sum of the differences between Kaplan-Meier survival curves and represented comparative data between two groups during a certain follow-up period. Data regarding multivariate HR and 95% CI were directly extracted from the studies. In cases where multivariate HR data were not available, univariate HR data were extracted. If both multivariate and univariate HRs were unavailable, the methodology recommended by Tierney et al was applied to reconstruct HR estimates and variations based on survival data read from Kaplan-Meier survival curves using the Engauge Digitizer (version 9.4). HR > 1 implies poorer survival in patients who underwent RFA than in patients who underwent surgical resection, whereas HR < 1 implies a survival benefit in patients who underwent RFA. RevMan version 5.3 (RevMan version 5.3; The Nordic Cochrane Center, The Cochrane Collaboration) and STATA version 12.0 (STATA Corp., College Station, TX) were employed for statistical analysis. Statistical heterogeneity was measured using the chi-squared Q test and I² statistic. Heterogeneity was considered to be > 50%. A fixed-effects model was used for meta-analysis when heterogeneity was not significant and a random-effects model was applied in case of significant heterogeneity. Begg and Egger tests were used to evaluate bias using STATA version 12.0. P values < .05 were considered statistically significant.

3. Results

3.1. Search results

The search process is presented in Figure 1. Searches were conducted on three databases, initially including 137, 1453, and 0 articles (1,590 in total) from Embase, PubMed, and Cochrane Library, respectively. After excluding duplicates and meeting summaries, 142 qualified full-text articles were left, 85 articles that did not meet the study design were removed, 6 studies did not describe liver cancer, 3 study introduced a case report, and there was no reliable data in 10 studies. Finally, 19 studies including 15,071 patients that met the conditions of the study, published between 2006 and 2017, were included in the meta-analysis (Fig. 1). All 19 studies reported at least one item of the applicable results.

3.2. Study characteristics

Almost all the studies were conducted in Asia (9 in China, 4 in Japan, 3 in Korea, 1 in France, and 2 in Italy). Three of the included studies were prospective studies and 16 were retrospective, and one study was prospective. In total, 7 studies analyzed RFS (including 1,945 patients), 6 analyzed DFS (including 838
| Study                     | Year | Country | Study design | Study period            | Follow-up duration | Median age (range), years | Study design | No. of patients | Child-Pugh Class (A/B/C) | End points | Sex (M/F) | Tumor Size (cm) (Range) |
|--------------------------|------|---------|--------------|-------------------------|--------------------|---------------------------|--------------|----------------|---------------------------|-------------|-----------|---------------------------|
| Min-Shan Chen et al. (2006) | 2006 | China | Pro          | 1999.11–2004.6         | 5 yr               | RFA Median (range): 51.9 (11.2) | RFA 71,     | NA             | HCC ≤5 cm                 | DFS OS RFA65/15 | ≥5 cm     |
| Abu et al. (2008)        | 2008 | Italy | Retro        | 1999-2003              | 5 yr               | RFA 67, Class A/B          | RFA 34,     | Child-Pugh A/B | DFS OS RFA27/7          | ≥5 cm       | ≥5 cm    |
| Ji-Wei Huang et al. (2010)| 2010 | China | Pro          | 2003.3–2005.1         | 5 yr               | RFA Mean (SD): 56.57 (14.30) | RFA 115,    | Class A/B | RFS OS RFA23/6 | HCC conforming to Milan criteria |
| Hiroki et al. (2011)     | 2011 | Japan | Retro        | 2004-2010              | 68.4 ± 8.7         | SR 65.9 (12.68)            | SR 115,     | RFS OS RFA65/67 | RFA 50/19, RFA 2.68 ± 0.49 | ≤5 cm       | ≤5 cm    |
| Hung-Hsu Hung et al. (2011)| 2011 | Taipei | Retro        | 2002-2007             | 60 mo              | RFA 16.4, Child-Pugh A/B | NA         | RFS OS RFA12/19 | ≤5 cm       | ≤5 cm       | ≤5 cm |
| Kai Feng et al. (2012)   | 2012 | China | Pro          | 2005.1–2008.3         | 3 yr               | SR 60.7 ± 12.45            | RFA 229,    | Child-Pugh A/B | SR 184/45, HCC 2 cm | ≤5 cm       | ≤5 cm    |
| Kiong-Ming Wong et al. (2012)| 2012 | China | Retro        | 2004-2009             | Median 36mo.       | RFA 36                     | Child-Pugh class A | RFS OS RFA18/18 | ≤5 cm       | ≤5 cm       | ≤5 cm |
| Jing-Hoang Wang et al. (2012)| 2012 | China | Retro        | RFA 2.5 years         | range 6–40 mo      | SR 85.1 ± 12               | NA         | RFS OS RFA18/18 | ≤5 cm       | ≤5 cm       | ≤5 cm |
| Kiyoshi Hasegawa et al. (2013)| 2013 | Japan | Retro        | 2000-2005             | 2.16 yr            | RFA 69 (52, 80)            | Child-Pugh A/B | OS RFA 3569/1979 | ≤5 cm       | ≤5 cm       | ≤5 cm |
| Kastunori Imai et al. (2013)| 2013 | Japan | Retro        | 2000.1–2011.4         | RFA 38.3 mo        | SR 66 (48, 77)             | Child-Pugh class A | OS DFS RFA23/6 | ≤5 cm       | ≤5 cm       | ≤5 cm |
| Maurizio Pompili et al. (2013)| 2013 | Italy | Retro        | 1999-2010             | RFA 41             | SR 86.8 (38–68)            | Child-Pugh class A | DFS RFA25/26 | ≤5 cm       | ≤5 cm       | ≤5 cm |
| Zhi-Peng Zhou et al. (2013)| 2015 | China | Retro        | 2003.7–2006.8         | SR 67 (35–69)      | SR 90                      | Child-Pugh class A | DFS RFA35/19 | ≤5 cm       | ≤5 cm       | ≤5 cm |
| Hyo-Joon Yang et al. (2014)| 2014 | Korea | Retro        | 2005.1–2006.12        | 8 yr               | SR 42.2 ± 7.6              | or B        | OS RFA9/20/0 | ≤5 cm       | ≤5 cm       | ≤5 cm |
| G.-A. Kim et al. (2015)  | 2015 | Korea | Retro        | 2000-2009             | RFA 61 (i.q.r. 40–82) | SR 55.7 ± 10.6             | RFS OS RFA26/71 | ≤5 cm       | ≤5 cm       | ≤5 cm       | ≤5 cm |

(continued)
| Study                          | Year | Country | study design | Study period            | Follow-up duration | Median age (range), years | No. of patients | Child-Pugh Class (A/B/C) | End points | Sex (M/F) | Tumor Size (cm) (Range) |
|-------------------------------|------|---------|--------------|-------------------------|--------------------|--------------------------|-----------------|--------------------------|-------------|-----------|------------------------|
| Arnaud Hocquet et al. (2015)  | 2015 | France  | Retro        | 2004.1–2013.12         | SR 66 (52–88)      | SR 54.4 (8–9)            | 273             | Child-Pugh class A       | OS          | RFA 18/20 or B       | SR 22 (30–28) |
| Po-Hong Liu et al. (2016)     | 2016 | China   | Retro        | 2002-2013              | RFA 60             | SR 60 (61–74)            | 103             | A or B                   | OS          | RFA 8/12              | SR 35 (25–40) |
| Tian-Pei Guan et al. (2016)   | 2016 | China   | Retro        | 2006.1–2010.5          | SR 43 mo           | SR 43 mo                | 60              | Child-Pugh A/B           | RFS OS      | SR 78/72              | RFA 2.0 (mean) |
| Kenichi Takayasu et al. (2017)| 2017 | Japan   | Pro          | 2000.1–2007.12         | 5 yr               | 67.0 ± 8.47              | 491             | Child-Pugh A/B           | RFS OS      | RFA 194/136           | 1.47 (0.36) |
| Seung-Ho Lee et al. (2017)    | 2017 | Korea   | Retro        | 2008-2010              | RFA 60 (38–88)     | SR 56 (33–76)            | 151             | Barcelona Clinic Liver Cancer (BCLC) stage A | DFS OS | SR 111/39 | SR 4.0 (3.1–5.0) |

DFS = disease-free survival, NA = not available, OS = overall survival, PFS = progression-free survival, Pro = prospective, RCT (randomized controlled trial), RFA = radiofrequency ablation surgical resection, RFS = recurrence/relapse free survival, Ret = retrospective, SR = surgical resection, TFS = tumor-free survival.
patients), 1 analyzed PFS (including 187 patients), 1 analyzed TFS (including 52 patients), and 14 analyzed OS (including 14,024 patients). The follow-up duration of 7 studies was ≥5 years. Tumor sizes in 10 studies were <3 cm, with 9 studies involving tumor sizes of 0 to 5 cm. There is a huge difference in the sample size of one study compared to the others. In fact, the study by Hasegawa et al (2013) included 74% of all patients evaluated overall in the 19 studies. Details of all studies were recorded, including the study period, follow-up duration, age, and number of patients (Table 1).

3.3. Literature quality evaluation
The quality of 19 studies was assessed according to the CRITICAL APPRAISAL OF PROGNOSTIC STUDIES (https://www.cebm.net/wp-content/uploads/2018/11/Prognosis.pdf; Fig. 2). Generally, the included studies were of high quality. Four of the studies were assessed as high risk and 3 as unclear risk of bias in the domain of defined representative sample measurements, since few were non-blinded or non-randomized. In the domain of prognostic factor, that is, follow-up time measurements, there was a high risk of bias in 2 studies and an unclear risk of bias in 3 studies owing to missing follow-up data or short follow-up time. The majority of studies were well described and monitored for adverse events based on objective criteria. In the domain of prognostic factor follow-up time measurements, there was a high risk of bias in 2 studies and an unclear risk of bias in 5 studies when monitoring for adverse events based on objective criteria.

3.4. Primary outcome: RFS
In total, 15 studies (3,323 patients) analyzed RFS as the primary endpoint. After combining HR, poorer RFS was predicted with RFA than with surgical resection. The random-effects model (HR = 1.55, 95% CI = 1.29–1.86, I^2 = 72.5%; Fig. 3) showed statistical significance and heterogeneity existed between studies. Potential publication biases were assessed using Funnel plots and
Figure 2. A Risk of bias graph: review author judgments on each risk of bias item presented as percentages across all included studies. B Risk of bias summary: review author judgments on individual risk of bias items for each included study.

Figure 3. Forest plots of HR for RFS and OS in relation to the patients treated with RFA and surgical resection (A, RFS; B, OS). Chi-squared test is a measurement of heterogeneity. $P < .05$ indicates significant heterogeneity. (Squares = individual study point estimates. Horizontal lines = 95% CI. Rhombus = summarized estimate and its 95% CI. Fixed: fixed-effects model.). ES indicates hazard ratio.
2 statistical (Begg and Egger) tests. Data from Funnel plots (Fig. 4) and Begg ($P = .373$) and Egger ($P = .342$) tests (Supplemental Figure 1, http://links.lww.com/MD/F622) indicated no significant publication biases. A sensitivity analysis was conducted to further estimate the impact on combined HRs (Fig. 5). Systematic exclusion of each study did not result in significant changes, suggesting that the results were stable.

Additional subgroup analyses were performed according to region, study design, tumor size, and endpoint (Table 2). Among the studies including RFS as an endpoint, 13 conducted in Asia had an HR of 1.54 (95% CI = 1.27–1.88, $I^2 = 75.3$%), and 2 conducted in Europe showed significant correlations (HR = 1.08, 95% CI = 1.08–2.11, $I^2 = 39.2$%). In terms of study design, the HR of 4 prospective studies was 1.52 (95% CI = 1.26–1.82, $I^2 = 31.6$%) and that of 11 retrospective studies was 1.58 (95% CI = 1.25–2.00, $I^2 = 77.2$%). According to the included tumor size, groups were divided into ≤3 and ≥5 cm subgroups. Meta-analyses illustrated that subgroups with tumor sizes of 3 cm had an HR of 1.56 (95% CI = 1.20–2.03, $I^2 = 77.2$%) and those with tumor sizes ≤5 cm had an HR of 1.55 (95% CI = 1.27–1.79, $I^2 = 39.2$%).
17.1%). Based on the endpoint, eligible studies were divided into RFS, DFS, PFS, and TFS groups. Subgroup analyses showed that the combined HRs of RFS and DFS were 1.64 (95% CI = 1.23–2.19, \(I^2 = 85.4\%\)) and 1.51 (95% CI = 1.27–1.81, \(I^2 = 24.0\%\)). No significant results were obtained for the PFS (HR = 1.36, 95% CI = 0.30–2.31) and TFS (HR = 0.97, 95% CI = 0.30–3.11) groups.

### 3.5. Primary outcome: OS

In total, 16 studies (17,025 patients) analyzed OS as the primary endpoint. After combining HR, poorer OS was predicted for RFA than for surgical resection. The random-effects model (HR = 1.61, 95% CI = 1.29–2.01, \(I^2 = 60.4\%\); Fig. 3) showed statistical significance and heterogeneity existed between studies. Potential publication biases were statistically assessed using Funnel plots, Beggs test and Egger test. Data from the Funnel plots (Fig. 4), Beggs test (\(P = .62\)), and Egger (\(P = .051\)) test (Supplemental Figure 1, http://links.lww.com/MD/F622) indicated no significant publication biases. A sensitivity analysis was conducted to further estimate the impact on combined HRs (Fig. 5). Systematic exclusion of each study did not result in significant changes, suggesting that the results were stable.

Additional subgroup analyses were performed according to region, study design, and tumor size (Table 2). Among the studies including OS as an endpoint, 13 conducted in Asia had an HR of 1.66 (95% CI = 1.40–1.97, \(I^2 = 12.6\%\)) and 3 in Europe had an HR of 1.67 (95% CI = 0.94–2.95, \(I^2 = 83.0\%\)). In terms of study design, the HR of 4 prospective studies was 1.37 (95% CI = 1.06–1.77, \(I^2 = 27.5\%\)) and that of 12 retrospective studies was 1.74 (95% CI = 1.30–2.33, \(I^2 = 67.3\%\)). Regarding the included tumor size, groups were subdivided into \(\leq 3\) cm and \(\leq 5\) cm subgroups. Meta-analyses illustrated that the subgroup with tumor sizes \(\leq 3\) cm had an HR of 1.13 (95% CI = 0.97–1.30, \(I^2 = 0.0\%\)), while those with tumors \(\leq 5\) cm had an HR of 1.89 (95% CI = 1.58–2.26, \(I^2 = 49.1\%\)).

### 4. Discussion

In clinical practice, to optimize treatment options for patients with small HCC, tumor characteristics, liver function reserve, and patient demographics are the key factors considered by clinicians or surgeons.\[18,31,32\] The clinical practice guidelines of the European Organization for Research and Treatment of Cancer recommend surgical resection as the first-line treatment for small HCC, with RFA considered the standard of care for patients who are not suitable for surgery.\[13,4\] However, this may sometimes be an extremely difficult decision. In recent years, RFA has been established as a valid alternative treatment for small HCC with comparable results to surgical resection in selected patients.\[13,31\] The potential merits of higher repeatability, greater tolerability, lower complication rates, and lower costs make RFA an attractive option.\[30\] Patients with comorbidities or poor liver functions are more likely to choose RFA because their comorbidities contraindicate SR. RFA is less invasive compared with SR, has a lower incidence, lower risk of complications, and is more likely to regenerate when relapsed.\[34,35\]

Focusing on long-term survival oncologic outcomes, meta-analysis results comprehensively and systematically showed that RFA is associated with poorer RFS (HR = 1.55, 95% CI = 1.29–1.86, \(I^2 = 72.5\%\)) and OS (HR = 1.61, 95% CI = 1.29–2.01, \(I^2 = 60.4\%\)) compared with SR. The evidence supporting this association was consistent in most subgroup analyses (region, study design, tumor size, and endpoints). In subgroup analysis performed according to region, RFS 1.54 (95% CI = 1.08–2.11, \(I^2 = 39.2\%\)) of the Europe and OS 1.66 (95% CI = 1.40–1.97, \(I^2 = 51.8\%\)).
of the Asian group showed statistical significance, and no statistical heterogeneity existed between studies. 13 studies analyzed the RFS of the Asian group, which revealed statistical significance, but statistical heterogeneity existed between studies. Three studies analyzed the OS of the European group, which showed no statistical significance between studies. The majority of studies were conducted in China, while only one and two studies on European subjects analyzed RFS and OS, respectively. This phenomenon is primarily affected by insufficient statistical power or is attributed to the higher prevalence of hepatitis B virus infection and HCC in China.[36] Further research on different regions is required to establish the prognostic value of RFA or surgical resection in patients with small HCC. All included patients met the Milan criteria (i.e., solitary HCC nodules ≤5 cm or up to 3 nodules, each < 3 cm). In subgroup analyses according to tumor size, both OS of 0 to 3 cm and 0 to 5 cm subgroups showed statistical significance, and no statistical heterogeneity existed between studies. A simulation study demonstrated a 3% increase in local recurrence rate after RFA, equivalent to a 1% increase in perioperative mortality.[37]

Thus, comparison of the treatment effects on the survival of patients with HCC should be performed separately for smaller (<3 cm) or intermediate (3 to 5 cm) nodules. In our meta-analysis, only Lee et al. (2017)[13] performed an analysis with RFS to compare the treatment effects on 3–5-cm nodules; however, HR showed no statistical significance (HR = 1.36, 95% CI = 0.8–2.312). We were unable to determine RFS and OS to compare the treatment effects on tumors 3 to 5 cm in size; therefore, a subgroup meta-analysis was not conducted. The difficulties in data collection limit further subgroup analysis with respect to the comparison of treatment effects on tumor sizes of 3 to 5 cm. Further randomized studies are warranted to explore the significance of 3 to 5-cm tumors on treatment-related survival of patients with HCC.

The choice of initial treatment for small HCC is largely dependent on the best interest of the individual patient and level of expertise in different institutions. Treatment algorithms to select patients who could benefit from SR or RFA are yet to be established. Randomized controlled trials provide high-level evidence by evaluating the clinical endpoints and using the most efficient and reliable method.[18] In subgroup analyses performed according to study design, both RFS (HR = 1.52, 95% CI = 1.26–1.82, I² = 31.6%) and OS (HR = 1.37, 95% CI = 1.06–1.77, I² = 27.5%) of the RCTs showed statistical significance, and no statistical heterogeneity existed between studies. In contrast, retrospective studies provide relatively low-level clinical evidence due to a potential selection bias. Both RFS (HR = 1.58, 95% CI = 1.25–2.00, I² = 77.2%) and OS (HR = 1.74, 95% CI = 1.30–2.33, I² = 67.3%) of retrospective subgroups showed statistical significance, though statistical heterogeneity existed between studies. Thus, data from both prospective and retrospective subgroups support our results.

In subgroup analyses performed according to endpoint, both RFS and DFS subgroups showed statistical significance between studies. Only 1 study analyzed PFS and one analyzed TFS, neither of which showed statistical significance. Further research on PFS or TFS should be conducted to establish the prognostic value of RFA and surgical resection in patients with small HCC.

Several previous meta-analyses[39–43] have compared the efficacy of RFA and surgical resection for improving the survival of patients with HCC through extraction of relative risk or values. To date, the results have been inconsistent. OR and relative risk are appropriate measures of dichotomous outcomes, since they simply involve the assessment of the number of events without consideration of the time when events develop. In the current study, HRs were extracted to compare the OS and RFS after surgical resection and RFA for HCC. The cumulative OS or RFS rates are time-to-event outcomes. Additionally, HRs are the most appropriate parameters to measure time-dependent outcomes.[17]

With the increasing application of laparoscopic and intraoperative ultrasound techniques, laparoscopic RFA (LRFA) presents a novel approach to local ablation. LRFA has other advantages, such as real-time security monitoring of the ablation process and accurate detection of tiny lesions, compared with percutaneous RFA. Gao et al (2016)[44] reported that LRFA led to fewer complications and lower hospitalization mortality compared with percutaneous RFA. However, limited information is available on its long-term efficacy. Microwave ablation (MWA) is also a common treatment for HCC. Several studies have found that MWA produced significantly larger zones than RFA.[45–47] Therefore, the apparent superiority of MWA over RFA has made it an alternative method to RFA, and it may be a more powerful technique than RFA. A meta-analysis showed that the efficacy of MWA in the treatment of HCC was similar to that of RFA.[48]

There was no significant difference between MWA and RFA in terms of prognosis and incidence of major complications. The long-term efficacy of LRFA and MWA in the treatment of small HCC was not clear. In addition, the efficacy of LRFA and MWA in the treatment of small HCC compared with that of surgical resection or RFA was not analyzed in this study. Multicenter, randomized, controlled studies are needed to verify the efficacy of LRFA and MWA against other therapies for small HCC.

The quality of the included studies is a limitation of our study. In addition, only published English language articles were included, and the exclusion of non-English articles may introduce language biases and erroneous conclusions. Further, the studies included in this meta-analysis (19 in total) mainly included Asian authors but also included two European works (one Italian and one French). Studies in this meta-analysis were mainly conducted in China, where the incidence of HCC is higher, which may be another potential cause of bias. Moreover, only published studies were included from electronic database searches; therefore, the possibility of publication biases cannot be excluded. However, it must be noted that evaluation of publication biases supported the reliability of our analysis. Furthermore, the Engauge Digitizer was used to extract HR data from survival curves indirectly, leading to potential imprecisions. Finally, selection biases could not be completely avoided. Selection of initial treatment was largely dependent on the best interest of the individual patient and level of expertise in different institutions. Treatment algorithms to identify patients who would benefit from surgical resection or RFA are yet to be established. Finally, the follow-up time of some studies was insufficient. With longer follow-up durations, disease relapse rates would likely be higher and prognostic performance would be different. Further multi-center and randomized controlled studies are necessary to confirm our findings.

5. Conclusion

In conclusion, surgical resection was associated with higher overall and disease-free survival than RFA. Thus, surgical resection remains the preferred treatment of choice for small HCC in cases where liver transplantation is not readily available.
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Conceptualization: Dongchun Xuan, Toufeng Jin.
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