Cost-Effectiveness of a Proteomic Test for Preterm Birth Prediction

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Background: Preterm birth (PTB) carries increased risk of short- and long-term health problems as well as higher healthcare costs. Current strategies using clinically accepted maternal risk factors (prior PTB, short cervix) can only identify a minority of singleton PTBs.

Objective: We modeled the cost-effectiveness of a risk-screening-and-treat strategy versus usual care for commercially insured pregnant US women without clinically accepted PTB risk factors. The risk-screening-and-treat strategy included use of a novel PTB prognostic blood test (PreTRM®) in the 19th–20th week of pregnancy, followed by treatment with a combined regimen of multi-component high-intensity-case-management and pharmacologic interventions for the remainder of the pregnancy for women assessed as higher-risk by the test, and usual care in women without higher risk.

Methods: We built a cost-effectiveness model using a combined decision-tree/Markov approach and a US payer perspective. We modeled 1-week cycles of pregnancy from week 19 to birth (preterm or term) and assessed costs throughout the pregnancy, and further to 12-months post-delivery in mothers and 30-months in infants. PTB rates and costs were based on >40,000 mothers and infants from the HealthCore Integrated Research Database® with birth events in 2016. Estimates of test performance, treatment effectiveness, and other model inputs were derived from published literature.

Results: In the base case, the risk-screening-and-treat strategy dominated usual care with an estimated 870 fewer PTBs (20% reduction) and $54 million less in total cost ($863 net savings per pregnant woman). Reductions were projected for neonatal intensive care admissions (10%), overall length-of-stay (7%), and births <32 weeks (33%). Treatment effectiveness had the strongest influence on cost-effectiveness estimates. The risk-screening-and-treat strategy remained dominant in the majority of probabilistic sensitivity analysis simulations and model scenarios.

Conclusion: Use of a novel prognostic test during pregnancy to identify women at risk of PTB combined with evidence-based treatment is estimated to reduce total costs while preventing PTBs and their consequences.

Keywords: preterm birth, cost effectiveness, progesterone, prognostic test

Plain Language Summary

- In the US, over 10% of infants are born prematurely (before 37 weeks’ gestation).
- Preterm birth (PTB) carries increased risk of short- and long-term health problems for infants, and higher healthcare costs for mothers and infants.
- Current strategies using clinically accepted maternal risk factors (prior PTB, short cervix) can only identify a minority of singleton PTBs.
- We created a decision-analytic model to compare a risk-screening-and-treat strategy to usual care in women at low risk of PTB by clinically accepted risk factors; the new
strategy consisted of a novel blood test (PreTRM®) that is
gnostic for PTB when applied in the 19th–20th week of
pregnancy, followed by treatment among women assessed
as higher-risk by the test with modelling of effect sizes
derived from the literature on multi-component high-inten-
sity case management and pharmologic interventions.

- The cost-effectiveness evaluation demonstrated that this
strategy of screening otherwise low-risk pregnant women
and treating those identified as higher-risk could lower total
healthcare costs and prevent preterm births.
- Sensitivity analyses confirmed robustness of the model and
showed the importance of treatment effectiveness in assess-
ing whether the risk-screening-and-treat strategy would be
cost-saving.

Introduction

Preterm birth (PTB) carries increased risk of short- and long-
term health problems for infants, and higher healthcare costs
for mothers and infants.1–3 In the US, about 10% of infants are
born prematurely (before 37 weeks’ gestation).4 The economic
burden of PTB has been estimated at $25 billion annually,
much of which is the direct cost of medical care in infancy.5,6
Within the category of PTB, risk of the most serious sequelae
and highest cost is not uniformly distributed, but much greater
for delivery at earlier weeks of gestational age.7,8 Each addi-
tional week in utero means better outcomes, including avoid-
ance of major neonatal morbidity and death.8

PTB is classified into subtypes of spontaneous PTB
(spontaneous onset of labor or following preterm premature
rupture of membranes; sPTB) and medically indicated PTB
(miPTB). PTB can follow preterm labor – often idiopathic or
related to occult intraamniotic infection, polyhydramnios,
uterine fibroids, and other predisposing or initiating condi-
tions. Typical medical indications include preeclampsia, pla-
cental abruption, intraterine growth restriction, and fetal
distress; in some cases, these conditions may also predispose
to spontaneous onset of labor.9

Although addressing the risk of prematurity has been
a national priority for decades, progress has been chal-
enged by many PTBs being spontaneous, without evident
clinically accepted maternal risk factors or fetal abnorma-
lities, and its heterogeneous and multifactorial nature.10
Reducing neonatal morbidity, mortality, and costs requires
identifying at-risk women sufficiently early to effectively
intervene. The shortcomings of clinically accepted mater-
nal risk factors (prior PTB, short cervix) which only iden-
tify a minority of singleton PTBs11–13 has led to interest in
the use of biomarker approaches for the identification of
otherwise at-risk women. A test to predict a majority of
preterm deliveries is optimally constructed with biomar-
kers that address general pathways downstream from mul-
tiple PTB etiologies.

We evaluated the cost-effectiveness of a risk-screen-
ning-and-treat strategy, compared to usual care, for
a population of commercially insured pregnant women in
the US without clinically accepted risk factors for PTB as
of the 19th-20th week of pregnancy. The approach
included one-time screening with a proteomic blood test
(PreTRM®, Sera Prognostics), which has demonstrated in
the 19th-20th week of pregnancy clinically valid predic-
tion of risk of sPTB <37 weeks and enrichment by its
biomarkers of early PTB (both spontaneous and medically
indicated) and adverse neonatal outcomes.14,15 In our
model, women identified as higher-risk via this test
would then be offered established PTB prevention mea-
sures, such as high-intensity case management (HICM)
and/or pharmologic interventions.16–18

One previous study evaluated the potential clinical and
cost outcomes of adopting a hypothetical test similar to
PreTRM using national data and found that if all pregnant
women in the US received the test, approximately 23,000
PTBs would have been prevented with savings of approxi-
mately $500 million in the first year of infant life.19 Our
study was designed to provide a complementary perspec-
tive using real-world data from a large US commercial
health plan, and to refine specific features of the previous
analysis. In particular, our objective was to develop
da detailed decision-analytic model to estimate the med-
ium-term cost-effectiveness of the PreTRM test compared
to usual care.

Methods

We conducted a cost-effectiveness analysis using pub-
lished literature and observational data from administrative
claims. Economic modelling best practices were
followed.20 Researchers’ access to claims data was limited
to data stripped of identifiers to ensure confidentiality.
HealthCore maintains data use agreements with the cov-
ered entities in compliance with the Health Insurance
Portability and Accountability Act. An Institutional
Review Board did not review the study since only this
limited data set was accessed.

Data Sources

The study population was based on a cohort of commer-
cially insured pregnant women age ≥18 with singleton
pregnancy, and a separate cohort of live-born singleton infants, based on birth events between January 1 and December 31, 2016. Data were extracted from the HealthCore Integrated Research Database (HIRD®), a repository of fully adjudicated, longitudinal claims from a large national commercial payer with membership in all 50 US states. Inclusion/exclusion criteria (Appendix Table A1) were further applied to ensure sufficient medical and pharmacy enrollment coverage and to select women corresponding to the test’s low-risk intended-use population; for example, women on progesterone in weeks 13 through 20 of gestation were excluded as this would be common routine care for women with prior PTB and/or short cervical length. Pregnancy, birth, neonatal outcome, and direct healthcare cost data were generated from the HIRD in weekly epochs of gestational age spanning from <23 weeks to ≥37 weeks (term).

Assumptions on test uptake and performance as well as treatment initiation and adherence were derived from published literature and expert opinion. Estimates of treatment effectiveness were based on published data examining the impact of different versions of HICM, progesterone, and bundles of these modalities on PTBs (a detailed derivation of the effectiveness estimates is provided in the Appendix). In our model, the treatment arm specifically included HICM consisting of up to two additional visits to a maternal-fetal medicine specialist, up to two additional transvaginal ultrasounds, and up to 10 additional nursing calls; daily low-dose (81mg) aspirin; and vaginal progesterone 200mg/day; all from weeks 22 up through 36. The model assumes similar treatment effectiveness in women determined to be high-risk by the PreTRM test as that observed for women determined to be higher-frequency use of progesterone, aspirin, or other treatments; our study did not further quantify routine practice patterns in this low-risk population but took these patterns and their associated PTB frequencies and cost outcomes as given. These women without clinically accepted risk factors would not have been subject to guidelines that generate particular routine interventions in current practice.

The model starts with a choice of usual care or risk-screening-and-treat at the 19th week of pregnancy. In the risk-screening-and-treat arm, chance nodes (governed by the probabilities listed in Table 1) determine whether the screening is actually performed, what the results are, whether the therapeutic treatment is initiated, and how adherent patients are to the treatment (Appendix Figure A1). At the end of each branch, identical Markov models advance in 1-week cycles until all pregnancies result in either a PTB or term birth (≥37 weeks) (Appendix Figure A2). We assumed the effect of the treatment began at week 23 to allow a 4-week time lag from the decision to test, including conducting the test and reporting results, to initiation of preventive measures and their earliest possible impact on PTB risk. Therefore, PTBs in weeks 19–22 are the same in both arms. Transition probabilities determine the chance of maintaining pregnancy to each subsequent week. Separate transition probabilities were derived for each arm. For usual care, the observed number of pregnancies and PTBs from the HIRD were used to calculate PTB risks. For the risk-screening-and-treat arm, the HIRD PTB risks were modified by the estimated weekly treatment effects, modelled as relative risk ratios, derived from previous literature on multi-component regimens (Appendix Table A2).

**Cost and Effectiveness Outcomes**

The model tracked direct healthcare costs (paid amounts as recorded in medical and pharmacy claims, aggregated across payer and patient components) over the medium term, defined as, for mothers, from beginning of pregnancy to 1 year post-delivery, and for infants from birth up to 30 months of life or disenrollment (to account for neonatal and early childhood developmental issues associated with PTB). All costs were adjusted to 2018 US dollars based on the Consumer Price Index for health care. Healthcare costs attributable to mothers included prenatal care costs (eg, ultrasounds, testing, treatments), delivery hospitalization costs, and post-natal costs, across all places of service. All-cause costs across all places of service, including neonatal care costs, were used for infants. Costs for progesterone and HICM were taken from public sources.
The primary effectiveness metric was the number of PTBs (births <37 weeks). The usual care arm distribution of PTBs (including both spontaneous and medically indicated) was identified from the neonates’ claims around the birth date using International Classification of Diseases, Tenth Revision, Clinical Modification (ICD-10-CM) codes (see Appendix “Methods Comments”). Other effectiveness outcomes included number of PTBs at <32 and <35 weeks, neonatal intensive care unit (NICU) admission, infant length of stay during delivery hospitalization (overall and in NICU), and neonatal morbidity/mortality index (NMI) changes. The NMI, a composite event score used in a multicenter trial of vaginal progesterone, was adapted to be estimated from claims (see Appendix “Methods Comments”).

Cost-effectiveness was assessed via the incremental cost-effectiveness ratio (ICER), expressed as the costs of adopting the risk-screening-and-test strategy minus costs for usual care, divided by the difference in total PTBs. Health outcomes and costs were not discounted given the short time horizon of the analysis.

Sensitivity Analysis
One-way and probabilistic sensitivity analyses (PSA) as well as scenario analyses were conducted to evaluate outcomes under variations in model parameters and to

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**Table 1 Model Inputs**

|                      | Base Case Value | Range (Lower–Upper) | Source     |
|----------------------|-----------------|---------------------|------------|
| Number of live births (Jan–Dec 2016) | 62,093          | NA                  | HIRD       |
| Number (%) of preterm births under usual care | 4360 (7.0%)     | NA                  | HIRD       |
| Proportion of women opting to be tested | 0.85            | 0.70–1.00           | Expert opinion |
| Characteristics of PreTRM® test |                |                     |            |
| Sensitivity*          | 0.75            | 0.5–0.83            | Saade 2016; data on file |
| Specificity*          | 0.74            | 0.49–0.82           | Saade 2016; data on file |
| Cost (one-time)       | $745            | $395–1000           | Expert opinion |
| Proportion of women identified as high-risk opting to receive treatment | 0.90            | 0.60–1.00           | Expert opinion |
| Characteristics of treatment for high-risk pregnancies | |                     |            |
| Proportion of treated women who adhere to the treatment | 0.80            | 0.60–1.00           | Expert opinion |
| Cost of high-intensity case management, including additional MFM specialist visits, weekly nurse monitoring, additional ultrasounds for cervical length assessment, and daily low-dose aspirin (per week) | $54.2           | $34.2–58.2         | CMS physician fee schedule for office visits and transvaginal ultrasounds (2019), plus generic cost of low-dose aspirin |
| Cost of progesterone (per week) | $26             | $20–50             | Average wholesale price for 200mg/day. Lower bound represents generic oral progesterone. Based on expert opinion, we assumed the majority of vaginal progesterone use is via suppository or direct insertion of an oral progesterone formulation.** |
| Effectiveness of treatment (risk ratio for preterm birth reduction) | 0.13–0.68; depending on gestational week | 0.01–0.83; depending on gestational week | Analysis of published literature on multi-component regimens (see Appendix Table A2 and “Methods Comments” section for details) |

**Notes:** Usual care assumes no testing and minimal progesterone use. *Sensitivity and specificity are linked via the following equation: Specificity = 2.0062 + 2.4210 × ln (1-Sensitivity) − 2.6145 × Sensitivity + 7.2006 × Sensitivity^2. **Based on expert opinion, we assumed the majority of vaginal progesterone use is via suppository or direct insertion of an oral progesterone formulation.

**Abbreviations:** CMS, Center for Medicare and Medicaid Services; HIRD, HealthCore Integrated Research Database; MFM, maternal-fetal medicine.
quantify the effects of uncertainty. Results from the one-way sensitivity analysis (using ranges described in Table 1 and Appendix Table A3) are presented as tornado plots. PSA used a Monte Carlo approach based on 10,000 randomly generated simulations of parameter values; results are presented as ICER scatter plots on the cost-effectiveness plane. Scenario analyses were conducted for key parameters not considered in the one-way and probabilistic sensitivity analyses such as treatment effectiveness (see Table 2 for details).

Results

Description of Study Cohort

We identified 62,093 live births in calendar year 2016 without claims indicating multiple gestation, and with at least one available indicator of gestational age (assessed on or within 30 days of birth). These 62,093 infants served as the basis of the model. Pregnant women aged ≥18, giving birth in an inpatient setting in 2016, were identified separately and included if they had no evidence of multiple gestation, early progesterone therapy (12–21 weeks of gestation), fetal abnormalities, preterm labor <21 weeks, or claims for stillbirth or abortion. Approximately two-thirds of women and infants were continuously enrolled in the health plan during the pregnancy and for at least 12 months after the birthdate. Mean (SD) age of mothers was 30.2 (5.1) years at the time of delivery. Women resided in all regions of the US, predominantly in the South (33%) and Midwest (28%). Sixty-eight percent of infants were delivered vaginally, and 52% were male.

The overall PTB frequency using infant claims was 7.0% in this selected population, which was used as the

| Table 2 Model Results: Base Case and Scenario Analyses |
|--------------------------------------------------------|
|                                   | Usual Care | Risk-Screening- | Incremental Difference | Savings per Prevented PTB* | Savings per Pregnant Woman |
|-----------------------------------|------------|-----------------|------------------------|---------------------------|---------------------------|
| **Base case**                     |            |                 |                        |                           |                           |
| Cost (USD)                        | $2611m     | $2558m          | -$53.6m                | $61,581                   | $863                      |
| Number of preterm births          | 4360       | 3490            | -870                   |                           |                           |
| **Scenario 1: Lower treatment effectiveness** |            |                 |                        |                           |                           |
| Cost (USD)                        | $2611m     | $2613m          | $1.7m                  | $3989 (ICER)              | N/A                       |
| Number of preterm births          | 4360       | 3924            | -436                   |                           |                           |
| **Scenario 2: Higher treatment effectiveness** |            |                 |                        |                           |                           |
| Cost (USD)                        | $2611m     | $2518m          | -$93.6m                | $70,722                   | $1508                     |
| Number of preterm births          | 4360       | 3036            | -1324                  |                           |                           |
| **Scenario 3: Alternate infant cost definition** |            |                 |                        |                           |                           |
| Cost (USD)                        | $2876m     | $2803m          | -$73.4m                | $84,344                   | $1182                     |
| Number of preterm births          | 4360       | 3490            | -870                   |                           |                           |
| **Scenario 4: Infant costs 0–12 months only** |            |                 |                        |                           |                           |
| Cost (USD)                        | $2438m     | $2387m          | -$51.1m                | $58,727                   | $823                      |
| Number of preterm births          | 4360       | 3490            | -870                   |                           |                           |
| **Scenario 5: Infant costs 0–12 months only, with alternate infant cost definition** |            |                 |                        |                           |                           |
| Cost (USD)                        | $2548m     | $2481m          | -$67.7m                | $77,779                   | $1090                     |
| Number of preterm births          | 4360       | 3490            | -870                   |                           |                           |

Notes: In Scenario 1, where incremental costs are positive, there are no savings and the number presented equals the actual ICER. Costs presented in 2018 USD. Incremental costs = risk-screening-and-treat strategy costs minus usual care costs. Risk-screening-and-treat is dominant (cost-saving and at least 1 preterm birth prevented) in the base case and all scenarios except Scenario 1. See Appendix for definitions of lower (Scenario 1) and higher (Scenario 2) treatment effectiveness. In Scenario 3, only costs from infants who were fully enrolled in their health plan over the 30-month follow-up time period were utilized (vs also using costs of those with <30-months enrollment; see Appendix Table A4). Scenario 4 truncated the modelling timeframe to the first 12 months after birth to observe a shorter-term impact (vs using 30 months). Scenario 5 used fully enrolled 12-month infant costs and a 12-month modelling horizon (a combination of Scenarios 3 and 4). *Equal to the absolute value of the ICER.

Abbreviations: PTB, preterm birth; m, million.
usual care rate for the economic analysis. As expected, costs by gestational age were higher on a per-patient basis for all preterm weeks (Figure 1 and Appendix Table A3). Births <32 weeks had the highest costs compared to term births, close to $400,000 on average for the year following birth. Maternal costs were more consistent across gestational ages than infant costs but still showed a relationship, with PTBs in weeks 23–36 averaging from 30% higher to more than double the cost of full-term births for the mother ($24,143 for week 37 and beyond).

Base Case Analysis
Using base case inputs of test uptake and performance, 15,570 mothers (out of 62,093) were identified as high-risk by the PreTRM test, with 11,210 receiving the treatment and remaining adherent. Under base case treatment effectiveness assumptions (see Appendix for details), 870 preterm births (20% of all preterm births in the cohort) were prevented in the risk-screening-and-test arm (Table 2). Incremental costs of the risk-screening-and-treat strategy in weeks leading up to delivery ($56 million) were more than offset by the reduction in costs of prevented PTBs ($109 million). Total costs for women (antenatal and 12-months postnatal) and infants (first 30-months) were therefore reduced by $54 million, an average net cost savings of $863 per pregnant woman, or $61,581 per prevented PTB. Risk-screening-and-treat is therefore a dominant strategy (improved outcomes at lower cost vs usual care) and no ICER was calculated. Fewer infants were born very early (33% reduction in gestational age <32 weeks; 26% reduction in <35 weeks) and fewer required admission to the NICU (9.8% reduction in NICU admissions). Length of hospital stay for infants was reduced by 7% overall and by 20% for NICU stays, and the proportion of infants with an NMI score ≥3, indicating severe morbidity or mortality, decreased by 23% (Table 3, Appendix Figure A3).

Sensitivity Analyses
Sensitivity analyses indicated that only the estimated effectiveness of the multi-component prevention regimen strongly influenced cost-effectiveness estimates (Figure 2A). Assuming lower effectiveness negated the cost savings

![Figure 1](https://doi.org/10.2147/CEOR.S325094) Health care costs for the first year of life and preterm birth rates by gestational age, usual care, 2016 data from the HealthCore Integrated Research Database.
gained in the base case. Nevertheless, cost per PTB prevented in this case ($3989; Table 2) remained favorable for decision-makers with a willingness-to-pay threshold of ≥$4000 per prevented PTB. When varying the cost, sensitivity, and specificity of the PreTRM test, eg, to account for any differences in the PTB risk of this commercially insured population vs the population used for the test validation, the risk-screening-and-treat strategy remained cost saving. Using PTB <32 or <35 as the endpoint, the risk-screening-and-treat strategy also remained cost saving in univariate sensitivity analysis (see Appendix Figures A4 and A5 for the ranking of influential parameters).

Scenario analysis tested changes to key parameters (higher/lower treatment effectiveness, alternative infant cost definition, use of a 12-month time horizon for infant costs) and found that the risk-screening-and-treat strategy remained dominant, resulting in both cost savings and avoided preterm births (Table 2). The only exception was in Scenario 1, which assumed lower treatment effectiveness, mirroring the one-way sensitivity analysis finding for this same variable.

The risk-screening-and-treat strategy was cost-saving in all base-case PSA simulations (Figure 2B). Since the base case one-way sensitivity analysis and Scenario 1 had identified low treatment effectiveness as potentially negating cost savings, we performed a separate PSA in Scenario 1 (where treatment effectiveness was set to “low”). Under this assumption, the risk-screening-and-treat strategy remained dominant (cost saving) in 32% of the 10,000 simulations, and was >80% likely to be cost-effective with a willingness to pay of ≥$30,000 per prevented PTB (Appendix Figure A6).

**Discussion**

Reducing PTBs is a leading goal of national health policy in the US. Women with a low-risk profile based on currently available information nonetheless carry considerable PTB burden, reflecting the need for better prognostic approaches. This study is the first economic evaluation of which we are aware using real-world data to assess the potential value of the PreTRM test for PTB risk in a commercially insured US population. Results suggest that the combination of this test with evidence-based multi-component prevention regimens in otherwise low-risk pregnant women provides a substantial reduction in PTBs while reducing overall health care spending. Fewer PTBs also translate into fewer sequelae for the newborns, which in our study can be seen in the lower rates

**Table 3 Model Results: Base Case Additional Endpoints**

| Outcome                                                                 | Usual Care | Risk-Screening-and-Treat Strategy | Incremental Difference (%) |
|------------------------------------------------------------------------|------------|-----------------------------------|----------------------------|
| **Total cost (USD) for births <37 weeks**                              | $566m      | $426m                             | ~$140m (~24.8%)            |
| **Number of PTBs (% relative to total population)**                     |            |                                   |                            |
| <32 weeks’ gestational age                                            | 525 (0.85%)| 351 (0.57%)                       | ~174 (~33.1%)             |
| <35 weeks’ gestational age                                            | 1698 (2.73%)| 1250 (2.01%)                      | ~448 (~26.4%)             |
| <37 weeks’ gestational age (total)                                     | 4360 (7.02%)| 3490 (5.62%)                      | ~870 (~20.0%)             |
| **Infant outcomes for births <37 weeks**                               |            |                                   |                            |
| Any NICU admission (% relative to total population)                    | 2821 (4.54%)| 2194 (3.53%)                      | ~628 (~22.2%)             |
| Length of stay (days)                                                  | 17.83      | 16.37                             | ~1.46 (~8.2%)             |
| NICU length of stay (days)                                            | 16.52      | 15.00                             | ~1.52 (~9.2%)             |
| Mean neonatal morbidity/mortality index over 30 days from birth        | 1.32       | 1.26                              | ~0.06 (~4.5%)             |
| Neonatal morbidity/mortality index ≥3 (% relative to total population) | 870 (1.40%)| 632 (1.02%)                       | ~238 (~27.4%)             |
| **Infant outcomes for all births**                                     |            |                                   |                            |
| Any NICU admission (% relative to total population)                    | 5939 (9.56%)| 5358 (8.63%)                      | ~581 (~9.8%)              |
| Length of stay (days)                                                  | 4.04       | 3.75                              | ~0.29 (~7.2%)             |
| NICU length of stay (days)                                            | 1.53       | 1.22                              | ~0.31 (~20.3%)            |
| Mean neonatal morbidity/mortality index over 30 days from birth        | 0.19       | 0.16                              | ~0.02 (~11.0%)            |
| Neonatal morbidity/mortality index ≥3 (% relative to total population) | 1037 (1.67%)| 801 (1.29%)                       | ~236 (~22.7%)             |

Notes: Costs presented in 2018 US dollars (USD). Percentages are relative to the total population of 62,093.

Abbreviations: NICU, neonatal intensive care unit; PTB, preterm birth; m, million.
of NICU admission, shorter NICU stays, and fewer newborns with high neonatal morbidity/mortality index scores. There is interest in a variety of other biomarker tests for PTB risk; however, uncertainties around their optimal clinical use have led to inconclusive findings with regard to cost-effectiveness.\textsuperscript{33,34}

The potential for cost savings from reducing the risk of PTB in the US is large. Based on our base case estimate of $863 in cost savings per pregnant woman without clinically accepted PTB risk factors, assuming approximately 4 million births per year of which \(~50\%)\ occur in commercial settings,\textsuperscript{35} and further assuming at least half have no evident

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**Figure 2** Model results: Sensitivity analyses. (A) Univariate sensitivity analysis (tornado diagram), base case. The tornado diagram (upper panel) ranks input parameters by their influence on the incremental cost-effectiveness ratio (ICER), from highest to lowest. (B) Probabilistic sensitivity analysis, base case. Each blue dot in the ICER scatter plot (lower panel) represents one of the 10,000 PSA simulation outcomes. The scatter plot is wedge-shaped with all mass in the second quadrant. All dots below the x-axis represent cost savings. In the base case, all simulations are associated with cost savings and PTB reductions.

**Abbreviations:** PTB, preterm birth; PSA, probabilistic sensitivity analysis; HICM, high-intensity case management; CI, confidence interval (95%).
risk factors for PTB, more than $850 million could be saved annually assuming full uptake of the risk-screen-and-treat strategy. Consistent with prior studies, healthcare expenditures in our cohort were inversely related to gestational age. Costs in our cohort tended to be higher than those in prior studies; this may be due to various methodological differences including the effects of medical care price inflation (the current study is among the most recent, using 2016 data inflation-adjusted to USD 2018) and our focus on commercially insured patients.

The PreTRM® test is a mass-spectrometry-based serum proteomics assay, which has strengths and limitations. The IBP4 (insulin-like growth factor-binding protein 4)/SHBG (sex hormone-binding globulin) protein ratio underlying the test shows connections to pathways potentially associated with PTB. Specifically, IBP4 regulates insulin-like growth factors, which are reported to play a role in nutrient delivery to the fetus. IBP4 is placentally expressed with higher circulating levels reported in women with growth restricted fetuses, suggesting that its increase may signal uteroplacental insufficiency. In addition, SHBG is a placentally expressed regulator of sex steroids, suppressed by proinflammatory mediators. Thus, reduction in SHBG may signal an infection or inflammatory state that can lead to increased availability of free steroids involved in pregnancy, some of which promote labor. Clinically meaningful prediction of PTB may be supported by biomarkers in general pathways such as placental dysfunction and inflammation which lie downstream of individual etiologies. More generally, mass-spectrometry-based serum proteomics draws strength from minimally invasive sample collection, a precise and robust measurement technology, and use of protein analytes that are produced as functional physiological signals. Multiplexing across pathways allows proteomics to address multifactorial conditions, while being somewhat limited by lack of full genome coverage, complexity, and, potentially, cost.

While the focus of clinical trials of pharmacologic interventions is usually on the reduction of sPTBs, the PreTRM test has been validated to predict sPTB <37 weeks, and the same biomarkers demonstrate an ability to enrich for early miPTBs across several potential medically indicated complications. Approximately 20–50% of PTBs are medically indicated. Claims data do not distinguish between spontaneous and medically indicated PTBs. Our analysis considers the average effect on amelioration of preterm birth to be independent of the reason for the preterm birth. One might consider this a limitation of the study from the standpoint of likelihood of pharmacologic interventions affecting medically indicated PTB. However, meta-analysis of vaginal progesterone treatment in high-risk women did not find evidence of differential effectiveness for amelioration of spontaneous vs medically indicated PTB. Furthermore, HICM studies do suggest specific benefit in reducing adverse outcomes associated with medically indicated PTB, which may in part be based on the intensive monitoring and behavioral modification aspects of such programs.

The model was populated with usual-care PTB and cost outcomes from a population of commercially insured real-world patients, selected in order to approximate the test’s intended-use population (singleton pregnancies lacking prior PTB or premature cervical shortening). We recognize that the inherent risk of PTB in the population will vary according to many genetic and environmental non-specific risk factors that are within intended use of the PreTRM test and included in the study population. We also note that the overall frequency of PTB (births <37 weeks) in our population of 7.0% is similar to or lower than the frequencies observed in the PreTRM clinical validation studies (7–11%), suggesting the test is suitable as baseline risks are aligned. Further, our sensitivity analysis considered the effect of different test performance and treatment effectiveness assumptions and indicated robustness of the base case results to such variations.

Strengths of this study include the large number of analyzed patients and generalizability to a US commercially insured population of pregnant women without clinically accepted risk factors for PTB. The distribution of births by gestational age matches prior published data. Differences in absolute frequencies may be due to our focus on commercially insured, singleton births. In general, conservative assumptions were used, including reliance on real-world data for treatment patterns and outcomes instead of assuming perfect adherence to guidelines. We assumed that multi-component treatment efficacy, as derived in our analysis from published study data, may be further attenuated by imperfect uptake and compliance. Our analysis used week-to-week estimates of cost and PTB risk, allowing us to estimate the benefits of even small-to-moderate changes in gestational age, which would not have been possible with a less granular model.

Limitations

Limitations of the study include reliance on treatment effectiveness estimates from published studies of multi-component HICM, progesterone, and novel bundled regimens in an evolving field. The interaction between effectiveness of this...
treatment and the population in which it will be used is dynamic, and to date the efficacy of progesterone has primarily been studied in women identified as high-risk using other prognostic methods, such as prior preterm birth or short cervix. While the PreTRM test would be applied to a population of otherwise low-risk women, the resulting treatment would only be targeted at those identified as higher-risk by the test, ie, a population where the treatment strategy is likely effective based on prior literature in other high-risk women. We did not model adverse events related to the risk-screening-and-test strategy; we note that safety evidence for progesterone is longstanding and reassuring. In light of the recent FDA proposal to withdraw the leading brand of injectable progesterone from the US market, our model focused on the vaginal formulations. Other treatments (eg, cerclage, antibiotics, and vitamin D) were not modelled and the conclusions from our model may not extend to their use.

The model has typical limitations common to the use of medical claims data, which lack clinical context and are subject to potential coding, billing, and recording errors as well as other inconsistencies. We did not model neonatal mortality as an isolated outcome; across the entire PTB population neonatal mortality is uncommon. Births <23 weeks’ gestation, stillbirth, or abortion were also not modelled. They occurred very rarely in this sample and are not known to be affected by the treatment. Patients without health insurance or those who are covered solely under public programs (eg, Medicaid) were not part of the analyzed population; commercially insured individuals are more likely to be currently employed and tend to have many factors that put them at lower risk than the remaining population, such as lower rates of smoking or sexually transmitted diseases. While we were not able to reliably ascertain race for the patients included in this dataset, it is also likely that fewer mothers were Black, Hispanic, or of other racial and ethnic minorities, compared to national statistics. In 2018, 63% of births to white mothers were covered by commercial insurance, versus 28% of births to Black mothers. Childbirth is an area of health care with particularly high disparities by race and economic status; consequently, the potential benefits of risk identification and treatment for patients underrepresented in the commercially insured population is an important area for future research.

Finally, it is important to note the limitations of our focus on direct health care cost. We did not consider quality of life (for mothers or infants) or long-term clinical sequelae, productivity, or other indirect costs, all of which may be substantially impacted by preterm birth. This is a conservative assumption likely to underestimate benefit of the risk-screening-and-treat strategy. Generating and incorporating appropriate data in this population would be an important task for future research.

Conclusions

In a health economic model, use of a novel blood test during the 19th–20th week of pregnancy to identify women at risk of PTB was shown to reduce health care costs and prevent delivery at early weeks of gestation in a representative population of commercially insured women in the US. These findings were consistent across a wide variety of possible scenarios in terms of test uptake, treatment adherence, treatment efficacy, and accrued costs.

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References

1. Barradas DT, Wasserman MP, Daniel-Robinson L, et al. Hospital utilization and costs among preterm infants by payer: Nationwide Inpatient Sample, 2009. *Matern Child Health J*. 2016;20(4):808–818. doi:10.1007/s10995-015-1911-y
2. Phibbs CS, Schmitt SK, Cooper M, et al. Birth hospitalization costs and days of care for mothers and neonates in California, 2009–2011. *J Pediatr*. 2019;204:118–125.e114. doi:10.1016/j.jpeds.2018.08.041
38. Truven Health Analytics. The cost of having a baby in the United States. In: Truven Health Analytics MarketScan® Study. 2013. Available from: https://www.nationalpartnership.org/our-work/resources/health-care/maternity/archive/the-cost-of-having-a-baby-in-the-us.pdf. Accessed June 2, 2021.

39. Forbes K, Westwood M. Maternal growth factor regulation of human placental development and fetal growth. J Endocrinol. 2010;207(1):1–16. doi:10.1677/JOE-10-0174

40. Crosley EJ, Dunk CE, Beristain AG, Christians JK. IGFBP-4 and −5 are expressed in first-trimester villi and differentially regulate the migration of HTR-8/SVneo cells. Reprod Biol Endocrinol. 2014;12(1):123. doi:10.1186/1477-7827-12-123

41. Qiu Q, Bell M, Lu X, et al. Significance of IGFBP-4 in the development of fetal growth restriction. J Clin Endocrinol Metab. 2012;97(8):E1429–E1439. doi:10.1210/jc.2011-2511

42. Larrea F, Díaz L, Caríño C, et al. Evidence that human placenta is a site of sex hormone-binding globulin gene expression. J Steroid Biochem Mol Biol. 1993;46(4):497–505. doi:10.1016/0960-0760(93)90104-5

43. Hammond GL. Diverse roles for sex hormone-binding globulin in reproduction. Biol Reprod. 2011;85(3):431–441. doi:10.1095/biolreprod.111.092593

44. Simó R, Sáez-López C, Beristain AG, Hernández C, Selva DM. Novel insights in SHBG regulation and clinical implications. Trends Endocrinol Metab. 2015;26(7):376–383. doi:10.1016/j.tem.2015.05.001

45. Bensimon A, Heck AJ, Aebersold R. Mass spectrometry-based proteomics and network biology. Annu Rev Biochem. 2012;81(1):379–405. doi:10.1146/annurev-biochem-072909-100424

46. Stout MJ, Busam R, Macones GA, Tuuli MG. Spontaneous and indicated preterm birth subtypes: interobserver agreement and accuracy of classification. Am J Obstet Gynecol. 2014;211(5):530.e531–530.e534. doi:10.1016/j.ajog.2014.05.023

47. Barton JR, Istwan NB, Rhea D, Collins A, Stanziano GJ. Cost-savings analysis of an outpatient management program for women with pregnancy-related hypertensive conditions. Dis Manag. 2006;9(4):236–241. doi:10.1089/dm.2006.9.236

48. Seegmiller RE, Nelson GW, Johnson CK. Evaluation of the teratogenic potential of delalutin (17 alpha-hydroxyprogesterone caproate) in mice. Teratology. 1983;28(2):201–208. doi:10.1002/tera.1420280208

49. Christian MS, Brent RL, Calda P. Embryo-fetal toxicity signals for 17alpha-hydroxyprogesterone caproate in high-risk pregnancies: a review of the non-clinical literature for embryo-fetal toxicity with progestins. J Matern Fetal Neonatal Med. 2007;20(2):89–112. doi:10.1080/14767050601178758

50. McNamara HC, Wood R, Chalmers J, et al. STOPPIT Baby Follow-up Study: the effect of prophylactic progesterone in twin pregnancy on childhood outcome. PLoS One. 2015;10(4):e0122341. doi:10.1371/journal.pone.0122341

51. Norman JE, Marlow N, Messow CM, et al. Vaginal progesterone prophylaxis for preterm birth (the OPPTIMUM Study): a multicentre, randomised, double-blind trial. Lancet. 2016;387(10033):2106–2116. doi:10.1016/S0140-6736(16)00350-0

52. Northen AT, Norman GS, Anderson K, et al. Follow-up of children exposed in utero to 17-alpha-hydroxyprogesterone caproate compared with placebo. Obstet Gynecol. 2007;110(4):865–872. doi:10.1097/01.AOG.0000281348.51499.bc

53. Food and Drug Administration. CDER proposes withdrawal of approval for Makena; 2020. Available from: https://www.fda.gov/drugs/drug-safety-and-availability/cder-proposes-withdrawal-approval-makena. Accessed June 2, 2021.

54. Crump C, Sundquist J, Howell EA, McLaughlin MA, Stroustrup A, Sundquist K. Pre-term delivery and risk of ischemic heart disease in women. J Am Coll Cardiol. 2020;76(1):57–67. doi:10.1016/j.jacc.2020.04.072

55. Crump C. Adult mortality after preterm birth-time to translate findings into clinical practice. JAMA Netw Open. 2021;4(1):e2033361. doi:10.1001/jamanetworkopen.2020.33361

56. Raju TNK, Pemberton VL, Saigal S, Blaisdell CJ, Moxey-Mims M, Buist S. Long-term healthcare outcomes of preterm birth: an executive summary of a conference sponsored by the National Institutes of Health. J Pediatr. 2017;181:309–318.e1. doi:10.1016/j.jpeds.2016.10.015

57. Wu P, Gulati M, Kwok CS, et al. Preterm delivery and future risk of maternal cardiovascular disease: a systematic review and meta-analysis. J Am Heart Assoc. 2018;7(2):e007809. doi:10.1161/JAHA.117.007809