INTRODUCTION

Neuropathic pain represents an area of largely unmet medical need, with significant impact on health-related quality of life (1). At present, neuropathic pain is treated with a variety of different medications with various mechanisms of action (2–7). Current treatments for neuropathic pain achieve some relief from the unrelenting pain, leaving significant room for improvement (8). An intriguing question is whether a greater effectiveness can be obtained by using rational combinations of drugs (9, 10). The optimization of a rational drug combination represents a major challenge (11). Pharmacokinetic-pharmacodynamic (PK-PD) modeling for the treatment of neuropathic pain. The focus of this study was to determine the influence of sildenafil on the pharmacokinetics (PK) of pregabalin with the objective of informing the design of a quantitative pharmacodynamic (PD) study.

Methods. The pharmacokinetics were determined in rats following 2-hr intravenous infusions of pregabalin at doses of 4 mg/kg/hr and 10 mg/kg/hr with and without a sildenafil bolus (2.2 mg) and steady state infusion (12 mg/kg/hr for 6 h). This PK model was utilized in a preclinical trial simulation with the aim of selecting the optimal sampling strategy to characterize the PK-PD profile in a future study. Eight logistically feasible PK sampling strategies were simulated in NONMEM and examined through trial simulation techniques.

Results. A two-compartment population PK model best described pregabalin pharmacokinetics. Significant model covariates included either a binary effect of sildenafil administration (30.2% decrease in clearance) or a concentration-dependent effect due to sildenafil’s active metabolite.

Conclusions. Analysis of simulations indicated that three post-PD samples had the best cost/benefit ratio by providing a significant increase in the precision (and minor improvement in bias) of both PK and PD parameters compared with no PK sampling.

KEY WORDS: neuropathic pain; optimal sampling; synergistic interaction; trial simulation.
mechanism of interaction is unknown, it is hypothesized that sildenafil may exhibit a PD synergistic interaction with pregabalin by increasing intracellular concentrations of cGMP. In turn, increased intracellular cGMP may amplify downstream regulation of the VGCC and modify pain perception in neuropathic pain.

The scope of this paper is to describe the development of a PK model identifying any potential PK interaction between pregabalin and sildenafil in rats with the objective of informing the design of a future quantitative PD study examining this interaction. Furthermore, the ultimate goal is the translation of this interaction to guide the first investigations of this drug combination in humans (29).

The overall process of informing the PD study design is outlined in Fig. 1. First, a PK study is conducted and a PK model developed using data collected from that study. This PK model provides the basis for investigating optimal PK sampling strategies for a PD study where PK sampling is not possible during the PD experiment. The necessity of designing the PK sampling in this way was demonstrated by a decrease in precision of PD endpoints during pilot studies where blood sampling was conducted in a chronic constriction injury rat model of neuropathic pain (Bender et al., unpublished). This distortion required that individual animal PK, ultimately for use in the PK-PD model, be estimated in animals without any blood samples taken during the PD portion of the experiment. To this end, a series of simulations was designed to evaluate the most parsimonious sampling strategy from among a series of eight sampling scenarios (17, 18, 30). The population approach to the PK study design allows the integrated application of PK samples from our dense PK study with the sparser sampling planned for our PD experiment. Through the application of nonlinear mixed effects modeling utilizing a Bayesian approach, the dense PK information can inform the PK model fit for those individuals with sparse sampling, allowing reconstruction of a full PK time course projected during the PD experiment (15, 31–34). Through examination of bias and precision of the prediction errors for the PK and PD parameters from each of the simulated sampling scenarios, the most parsimonious sampling strategy was selected.

**MATERIALS AND METHODS**

**Chemicals**

Pregabalin, sildenafil and the N-methyl metabolite of sildenafil were obtained from the Compound Control Unit of Pfizer Global Research and Development, Pfizer UK.

**Animals**

Male Sprague-Dawley rats (200–250 g), obtained from Charles River Laboratories (Margate, UK), were housed under a 12:12 h light-dark cycle with food and water *ad libitum*.

**Surgical Procedures**

All experiments were conducted in a chronic constriction rat model of neuropathic pain as previously described by Bennett and Xie (35). Briefly, this surgical procedure involved tying four loose ligatures around the right sciatic nerve. After this procedure, the animal developed a peripheral mononeuropathy which resembles the human condition in its response to static and dynamic allodynia (35).

Jugular venous catheters (JVC) and carotid artery catheters (CAC) were implanted according to standard surgical procedures. Briefly, this entailed making a small incision in the skin over the right jugular vein or left carotid artery and blunt dissection through the subcutaneous fat and connective tissue to expose the vessel. A section of vessel approximately 5 mm in length was then isolated, a loose ligature was tied caudally and the cranial end was ligated. A small incision was made between these ligatures, and the catheter was introduced and secured in place with the preplaced sutures. The catheter was then tunnelled subcutaneously to exit through the scapula region into a button of a Dilab® automated blood sampling apparatus. A local ethics review committee approved all experimental protocols.

**Drug Administration and Blood Sampling**

During this investigation, the PK of pregabalin was determined in the presence of either a steady-state concentration of sildenafil, targeted at maximum target inhibition, or placebo. Drugs were administered as an intravenous infusion through the jugular venous catheter. Volumes and infusion rates were kept constant throughout all treatment groups. Sildenafil or placebo was administered by a combined bolus and continued infusion. Specifically, a 1 mL/kg bolus containing 2 mg/kg sildenafil was administered over 10 mins followed by an infusion rate of 2 mL/kg/hr containing 12 mg/kg/hr of sildenafil for the duration of the planned 6 h PD experiment. Pregabalin was administered as a 2-hr infusion (doses of 10 mg/kg/hr or 4 mg/kg/hr), as this was determined during pilot studies to be the optimal administration schedule. Specifically, this infusion length allowed characterization of the PD endpoints during both the upswing and downswing in concentration change during the course of the 6-hr PD sampling planned for the

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**Fig. 1.** Flowchart depicting PK modeling and the simulations conducted in this study.
Pharmacokinetic Study Design

The PK study was designed as a crossover study on two occasions with a washout between occasions as outlined in Table I. A total of 28 rats divided evenly \((n=7)\) among four treatment groups were evaluated. The purpose of this multi-day study was to allow independent estimation of the intra-individual and the intra-occasion variability as well as to conserve animals.

PK sampling was conducted in a staggered population approach using seven different sampling schedules corresponding to one rat from each of the four treatment groups (Table II). Up to eight PK samples were taken from each animal between 0.33–22 h after the start of the infusion on a particular experimental day (Table II). Each rat was sampled on day 1 and with at least a three-day washout period (day 4). An optimal sampling algorithm (D-optimal) was utilized as implemented in the ADAPT II software (36) to aid in the selection of these time points.

Table II. Sample Times for Individual Rats During the PK Study. Animal Number 1–7 Corresponds to Rat 1–7 in Each of the Four Treatment Groups

| Animal number | Sample times (hrs) |
|---------------|--------------------|
| 1             | 0.33, 1.33, 2.33, 3.5, 5.5, 7.5, 12, 22 |
| 2             | 0.33, 1.33, 2.33, 4.5, 8.5, 12.5, 17, 22 |
| 3             | 0.66, 1.66, 2.5, 4, 8, 11.5, 16, 22 |
| 4             | 0.66, 1.66, 2.5, 5, 9, 11, 15, 22 |
| 5             | 1, 2, 3, 4.5, 8, 12, 16, 22 |
| 6             | 1, 2, 3, 5, 9, 12.5, 17, 22 |
| 7             | 1, 2, 3.5, 5.5, 9.5, 12.5, 16, 22 |

Drug Analysis

Plasma samples were analyzed for quantification of pregabalin, sildenafil and the major active N-methyl metabolite of sildenafil by an LC-MS/MS method briefly described below.

The extraction of pregabalin, sildenafil and the N-methyl metabolite was performed using solid phase extraction. Rat plasma samples, standards and quality controls (50 μL) were mixed with 450 μL water and then 500 μL 1 M mono chloroacetic acid in methanol/water (1:9 v/v) prior to extraction using activated (2 ml methanol) solid-phase extraction sorbent (Oasis MCX, Waters Corporation) packed in a 96 well format with 30 mg sorbent per 2 ml well. Following sample application, the cartridges were washed sequentially with 1 ml water followed by 1 ml of 1 M mono chloroacetic acid in methanol/water (1:9 v/v) followed by 1 ml methanol. Pregabalin, sildenafil and the N-methyl metabolite and internal standards (UK-464,243 and UK-343,664) were eluted with 1 ml of ammonium hydroxide:methanol (5:95 v/v) into 2 ml deep, 96 well blocks. These extracts were retained in the blocks, evaporated to dryness under a nitrogen stream (37°C) and resuspended in 200 μL 2 mM ammonium acetate, pH 4.0, in methanol water (10:90 v/v).

Samples were analyzed on a Sciex API 3,000 mass spectrometer in positive mode using turbo ionspray. A fast gradient HPLC was used with a Chromolith Speed Rod RP-18e, 50×4.6 mm, 5 μm (Merck) at 3 ml/min with flow splitting ratio 1 to 5. Mobile phases of 0.027% HCOOH v/v + 10 mM ammonium acetate 90:10 water-CH₃OH (organic mobile phase) and 0.027% HCOOH v/v + 10 mM ammonium acetate 90:10 CH₃OH:water (aqueous mobile phase) were used and prepared fresh on each day. The method scanned for all compounds simultaneously. The MRM transitions used for each compound were as follows: pregabalin (160 to 142); UK-464,242 (184 to 170); sildenafil (475 to 58); the N-methyl metabolite of sildenafil (461 to 283); and UK-343,664 (566 to 346).

Intra- and inter-day assay accuracy and precision were assessed for each compound at 150, 2,000, 9,000 ng/ml for pregabalin and 3,500, 1,800 ng/ml for sildenafil and UK-103,302 spanning the calibration range \((n=5)\). Accuracy was assessed based upon the percentage of the ratio of the found concentration on the nominal concentration. Intra-day assay accuracies ranged from 95.5% to 107.6% for pregabalin,
101.2% to 111.1% for sildenafil and 99.8% to 107.2% for N-methyl metabolite of sildenafil. Inter-day (3 days) assay accuracies ranged from 101.2% to 108.3% for pregabalin, 101.2% to 108.3% for sildenafil and 100.0% to 102.8% for UK-103,320. Precision of the assay was expressed as %CV (coefficient of variability) calculated from five replicates at three concentrations. %CV for intra-day assay was 10.1 (150 ng/ml), 7.3 (2,000 ng/ml), 5.5 (9,000 ng/ml) for pregabalin, 10.7 (3 ng/ml), 5.1 (500 ng/ml), 9.6 (1,800 ng/ml) for sildenafil, and 11.3 (3 ng/ml), 5.4 (500 ng/ml), 8.4 (1,800 ng/ml) for UK-103,320. %CV for inter-day assay was 13.2 (150 ng/ml), 10.2 (2,000 ng/ml), 11.6 (9,000 ng/ml) for pregabalin, 14.5 (3 ng/ml), 5.3 (500 ng/ml), 7.8 (1,800 ng/ml) for sildenafil, and 13.8 (3 ng/ml), 6.6 (500 ng/ml), 9.0 (1,800 ng/ml) for UK-103,320. Quantitation was linear over the following concentration ranges: pregabalin (100–1,800 ng/ml), sildenafil (3–2,000 ng/ml) and the N-methyl metabolite of sildenafil (5–2,000 ng/ml). The lower limit of quantitation for pregabalin, sildenafil, and the N-methyl metabolite of sildenafil was 70 ng/ml, 1 ng/ml, and 1 ng/ml, respectively.

Data Analysis

All PK modeling was accomplished utilizing a standard nonlinear mixed effects approach implemented within NONMEM (double precision Version 5, level 1.1, GloboMax, Hanover, MD) (37, 38). The models consisted of a structural model that described the disposition of the drug following IV administration, and a pharmacostatistical model that described the between-subject, between-occasion and residual variability. Diagnostic graphics, exploratory analyses, and post-processing of NONMEM outputs were performed using S-PLUS (Version 7.0 for windows, Insightful, Seattle, WA) with a NONMEM interface developed by LAP&P consultants BV (Leiden, the Netherlands).

The first order conditional estimation method (FOCE-INTER) was used for model building. The adequacy of each model was assessed using standard goodness-of-fit plots, the precision of the parameter estimates (as calculated using the COVARIANCE option of NONMEM) and the NONMEM objective function value (OFV). When discriminating between hierarchical models based on OFV, the likelihood ratio test was applied. This test is based on the property that the ratio of the NONMEM objective function values (OFV) was asymptotically $\chi^2$ distributed. An objective function value decrease of 10.8 units was considered significant ($\chi^2$ $p<0.001$ df =1).

Standard errors for all parameters were obtained using the covariance option in NONMEM. The population PK analysis was performed using the NONMEM subroutine ADVAN 3 TRANS 4 (two-compartment model). It was assumed that both the between-occasion variability (BOV) and the between-subject variability (BSV) of the PK parameters were log-normally distributed. Therefore, the relationship between a PK parameter ($P$) and its variance was expressed as:

$$P_i = P_{TV} \times e^{\varepsilon BSV_i + \eta BOV_i}$$

(1)

Here, $P_i$ was the value of PK parameter for the $i$th individual, $P_{TV}$ was the typical value of $P$ for the population, and $(\eta BSV_i, \eta BOV_i)$ denoted the difference between $P_i$ and $P_{TV}$, independently, which was identically distributed with a mean of zero and variance of $\sigma^2$. $\eta BSV$ refers to the between-subject variability, while $\eta BOV$ is the between-occasion variability.

The residual variability was examined using additive, proportional, and combined error structures as described below:

$$\text{Additive error: } y_{ij} = \hat{y}_{ij} + \varepsilon_{ij}$$

(2)

$$\text{Proportional error: } y_{ij} = \hat{y}_{ij}(1 + \varepsilon_{ij})$$

(3)

$$\text{Combined additive and proportional error: } y_{ij} = \hat{y}_{ij}(1 + \varepsilon_{ij}) + \varepsilon_{ij}'$$

(4)

Here, $y_{ij}$ was the $j$th observation in the $i$th individual, $\hat{y}_{ij}$ was the corresponding model prediction, and $\varepsilon_{ij}$ (or $\varepsilon_{ij}'$) was a normally distributed random error with a mean of zero and a variance of $\sigma^2$.

The final model was developed by testing the effect of subject-specific covariates’ bodyweight, age, time post CCI-surgery, time post catheterization surgery, sildenafil concentration, and sildenafil’s metabolite concentration. All covariates were initially modeled as continuous. Sildenafil was also modeled as a discrete covariate as the steady state infusions used during the PK study resulted in a relatively stable concentration of sildenafil which would saturate its target over the experimental period of interest. Stepwise covariate selection was utilized for the covariate model-building (39–43). First, exploratory covariate selection was performed by examination of the normalized eta deviation between individual post-hoc parameter estimates and candidate covariates. Subsequently, various forms of parameterization of the selected covariates were added to the base model and evaluated for significance by observing $\Delta$OFV and diagnostic plots. Only the single covariate parameterization producing the most significant increase in goodness of fit then moved on to the next stage. This continued until no significant improvements in model fit could be gained through further covariate inclusion.

The following example shows the effect of a continuous covariate on CL:

$$TVCL = \theta_{CL} \times (\text{Cov/MedCov})^{\theta_{cov}}$$

(5)

$$CL = TVCL \times \exp(\eta_i)$$

(6)

$TVCL$ was the typical value for the population; $\eta_i$ was the random effect representing the difference of the $i$th subject from the population mean. The random effects of between-subject variability were assumed log-normally distributed, with a mean of zero and standard deviation of $\sigma$. $\text{Cov}$ was the continuous covariate that was affecting $CL$, and $\text{MedCov}$ was the median $\text{Cov}$. The following example shows the effect of a discrete covariate sildenafil presence (SLDB) on $CL$:

$$TVCL = \theta_{CL} + (1 − SLDB) \times \theta_{SLD}$$

(7)

When SLDB was 0 (sildenafil absence = 0, sildenafil presence = 1), TVCL equals $\theta_{CL}$ and when SLDB was 1, the
\( \theta_{\text{SLD}} \) term was subtracted from the population estimate of CL. The final model was evaluated utilizing bootstrapping with resampling technique to examine the stability of the final model and estimate confidence intervals of the parameters (44–46). This technique consists of fitting the model to 100 bootstrapped replicates of the dataset using the bootstrap option in the LAP&P software interface for SPLUS. Parameter estimates for each of the replicate data sets were obtained. The results from 500 successful runs were obtained, and the mean and 2.5th and 97.5th percentiles (denoting the 95% confidence interval) for the population parameters were determined and compared with the estimates of the original data.

A posterior predictive check was performed (47). A posterior predictive check involves utilizing the simulation option within NONMEM to create 90% prediction bands. These are obtained by simulating 1,000 replications using the final model and a dataset including identical experimental design. For each observation, the values of the 51st highest and the 51st lowest prediction become the edges of the 90% prediction band. Observed values are then plotted on the same scale as these bands to help visualize the adequacy of model fit. After completion of these steps, the final descriptive PK-PD model was proposed.

Simulations to Evaluate Pharmacokinetic Sampling Paradigms

The population PK model was used as the basis of simulations designed to identify the most parsimonious PK sampling schedule for implementation in the future PK-PD study. The simulation study design was selected to mimic the pregabalin-only portion of the planned PD study. Doses and number of animals per group were previously determined from pilot experiments based on the nature of the PD interaction and what could be realistically utilized in these types of experiments, respectively. As such, the motivation for the simulation PK study was intended only to optimize the PK sampling portion of future PD experiments.

PK profiles for pregabalin doses of 4 and 10 mg/kg/hr were simulated with 12 rats in each dose group for each sampling scenario. Simulated samples were generated from these profiles according to the sampling schedule. A total of eight sampling scenarios were examined as outlined in Table III. Next, the simulated samples were combined back with the real-world PK samples to serve as an anchor for the population PK model. These new datasets were then remodeled, and the newly predicted individual PK parameters were compared with those used to create the simulation. Prediction errors were then calculated to determine the bias and precision of predictions provided through each sampling scenario. After this evaluation process was completed for the PK model, a PD model examining the impact of the various PK sampling schedules on the ability to predict PD was evaluated in a similar fashion. More detailed methods are described below.

Sampling scenarios to be evaluated (Table III) were selected to compare various post-PD sampling schedules with the best case (eight PK samples) and worst case (no samples) scenarios. In addition, a sampling design with only a single PK sample taken during the PD portion of the experiment was simulated. The first four scenarios included only post-PD sampling. The fifth scenario examined the effects of dense sampling during the entire experiment, but on another occasion. The sixth schedule examined the hypothetical best-case scenario of dense sampling during the entire experiment (were it possible). The seventh scenario examined the potential benefit of including one sample during the PD experiment and one immediately after. Finally, the worst case scenario of obtaining no individual animal PK was assessed.

Monte-Carlo simulations were used to examine the impact of the various sampling schedules on the identification of the individual animal-specific PK. In this procedure the simulation function of NONMEM creates “simulated rats.” These simulated rats have PK parameters (simulated parameters) selected from within the distribution of possible parameters defined by the final PK model with the assistance of a random number generator. Simulated PK profiles were then generated for each of the simulated rats with samples taken from each of these profiles according to the proposed sampling schedule. These simulated concentration points were then combined with the complete PK dataset (real dataset from the dense PK study), along with the “true” individual-specific simulated parameter estimates (for future comparison). This new simulated dataset was modeled in NONMEM utilizing the same methods as described in the PK analysis section generating predicted parameters. This whole process is outlined in Fig. 1.

This cycle of simulation and estimation was repeated 100 times for each of the eight sampling schedules creating 100 “simulated experiments” for each. Model performance was then assessed examining how close the predicted parameters were to the “true” or simulated parameters using the prediction error on the individual clearance parameter estimates. The prediction error (\(PE\)) was calculated as follows:

\[
PE = \frac{\theta_{\text{pred}} - \theta_{\text{sim}}}{\theta_{\text{sim}}} \times 100\%
\]

The prediction errors were visualized for each sampling schedule using boxplots to examine their bias (median prediction error) as well as precision (quartiles) for each replication. The bias and precision along with practical considerations served to guide the sampling protocol selection.

A previously developed PK-PD model was then applied to examine the ability to accurately identify the EC50 for static allodynia given three sampling scenarios (best(6), worst...
In this PK-PD model, the PK approach described above was applied. This was coupled with an effect compartment, which was used to drive the PD model. The PD model consisted of a sigmoid \( E_{\text{MAX}} \) model to relate the concentration of drug available at the effect site \( (C_e(t)) \) to the observed response expressed as a percentage change in response to static allodynia from baseline. In this model, \( E_{\text{MAX}} \) is the maximum change in response the drug can produce (fixed to 1 or 100%), \( EC_{50} \) is the value of \( C_e(t) \) producing 50% of the \( E_{\text{MAX}} \) value, and Hill influences the steepness of the relationship:

\[
E(t) = \frac{EMAX \cdot C_e(t)^{Hill}}{EC50^{Hill} + C_e(t)^{Hill}}
\]

Parameters for the \( E_{\text{MAX}} \) model were obtained from a pilot study in a chronic constriction injury model of neuropathic pain, using the difference in paw withdrawal threshold upon stimulation with von Frey hairs as a pharmacodynamic endpoint (28). The study was conducted in male Sprague Dawley rats (Charles River Laboratories (Margate, UK)). The values of the (population) pharmacodynamic parameters were \( Keo = 6.27 \) h (23.6%), \( EC_{50} = 9.36 \) ug/mL (5.3%), \( E_{\text{MAX}} = 1 \), and Hill = 3.8 (8.3%), assuming a normal distribution. PD parameters and plots were simulated and re-evaluated in NONMEM in the same way as the PK simulations. From the simulated and predicted PD parameters, prediction errors for \( EC_{50} \) were calculated and evaluated in the same way as the PK parameters and considered in the selection of the most appropriate sampling strategy.

**RESULTS**

Concentrations of pregabalin, sildenafil and the active hydroxyl metabolite of sildenafil are displayed in Fig. 2. Maximal concentrations of pregabalin were reached at the end of the 2-hr infusion and are roughly 22,000 and 10,000 ng/mL respectively for the 10 mg/kg/hr and 4 mg/kg/hr pregabalin infusions. When sildenafil was administered concomitantly, visible increases in the concentrations of pregabalin can be observed.

A two-compartment model was chosen as the final structural model to describe the concentration time course of pregabalin. Estimated model parameters are listed in Table IV. Both between-subject variability and between-occasion variability were determined in the base model for clearance (BSV 0.024(24.7%), BOV 0.098(27.1%)) and the volume of the central compartment (BSV 0.017(9.4%), BOV 0.015(9.0%)). A proportional error model was selected to account for the residual variability. Overall diagnostic plots demonstrated a good model fit (Fig. 3). Observed values versus both population and individual predicted were near the line of identity. Weighted residuals were relatively small in magnitude and randomly distributed when viewed versus time and concentration.

Of all the covariates investigated, only sildenafil and its metabolite had a significant impact on the model. This is likely due to the homogeneous nature of the inbred rat strain and controlled conditions of the experiment. Two final covariate models incorporating the presence of sildenafil were decided upon. First, a binary model was implemented...
Table IV. Objective Function Values (OFV) and Parameter Value Estimates for the Two-Compartment Pregabalin PK Model. Sildenafil Interaction was Incorporated as Either a Binary $(1 - \text{SLDB})\theta_{\text{SLD}}$ or Continuous $(1 - [\text{SLDM}] / (\theta_{\text{SLD}} + [\text{SLDM}]))$ Covariate Effect on Clearance. $\omega_1$ and $\omega_2$ Denote Between-Subject and Between-Occasion Variability, Respectively, for the Listed Model Parameter

| OFV                                | Original PK Model | Binary Sildenafil Covariate | Continuous Sildenafil Metabolite Covariate |
|------------------------------------|-------------------|-----------------------------|-------------------------------------------|
|                                    | Value             | CV%                         | LLCI | ULCI | Value             | CV%                         | LLCI | ULCI | Value             | CV%                         | LLCI | ULCI |
| CL (L/hr)                          | 0.034             | 19.08%                      | 0.021 | 0.046 | 0.052             | 14.59%                      | 0.037 | 0.066 | 0.051             | 8.83%                      | 0.042 | 0.060 |
| V1 (L)                             | 0.270             | 3.38%                       | 0.252 | 0.288 | 0.271             | 3.38%                       | 0.253 | 0.289 | 0.272             | 3.27%                       | 0.255 | 0.289 |
| Q (L/hr)                           | 0.0225            | 32.00%                      | 0.0084 | 0.0366 | 0.0151            | 35.76%                      | 0.0045 | 0.0257 | 0.0229            | 18.08%                      | 0.0148 | 0.0310 |
| V2 (L)                             | 6.75              | 41.48%                      | 1.26  | 12.24 | 2.77              | 51.99%                      | −0.05 | 5.59  | 2.45              | 38.82%                      | 0.59  | 4.31  |
| $\alpha_1$CL                      | 0.024             | 24.74%                      | 0.012 | 0.035 | 0.017             | 15.91%                      | 0.012 | 0.022 | 0.013             | 14.63%                      | 0.009 | 0.016 |
| $\omega_1$V1                      | 0.017             | 9.38%                       | 0.014 | 0.020 | 0.018             | 9.26%                       | 0.014 | 0.021 | 0.016             | 8.72%                       | 0.013 | 0.018 |
| $\alpha_1$Q                       | 0.430             | 40.00%                      | 0.093 | 0.767 | 0.439             | 40.87%                      | 0.087 | 0.791 | 0.439             | 15.94%                      | 0.055 | 0.106 |
| $\omega_2$CL                      | 0.098             | 27.07%                      | 0.046 | 0.151 | 0.061             | 19.52%                      | 0.038 | 0.085 | 0.071             | 18.52%                      | 0.045 | 0.097 |
| $\omega_2$V1                      | 0.015             | 8.96%                       | 0.013 | 0.018 | 0.015             | 8.75%                       | 0.012 | 0.017 | 0.016             | 8.33%                       | 0.013 | 0.019 |
| $\theta_{\text{SLD}}$             | 0.302             | 23.05%                      | 0.166 | 0.438 | 0.302             | 23.05%                      | 0.166 | 0.438 | 1350              | 28.37%                      | 599   | 2101  |
| $\sigma_{\text{residual}}$        | 0.029             | 18.32%                      | 0.019 | 0.040 | 0.029             | 18.53%                      | 0.019 | 0.040 | 0.029             | 18.85%                      | 0.018 | 0.039 |

Fig. 3. Diagnostic plots for the pregabalin PK model: A observed vs. individual predicted, B observed vs. population predicted, C weighted residuals vs. time, D and weighted residuals vs. predicted pregabalin concentration. Symbols depict data points from rats administered 4 mg/kg without sildenafil (circles), 10 mg/kg without sildenafil (squares), 4 mg/kg with sildenafil (+), and 10 mg/kg with sildenafil (triangles).
examining the presence or absence of sildenafil in the experiment. The final parameters for this model are shown in Table IV. The effect on clearance was demonstrated graphically in Fig. 4. Sildenafil resulted in a reduction in pregabalin clearance of 30.2% with a CV of 23.1%. The optimal descriptive model fit utilizing the continuous concentrations of sildenafil and its metabolite was also sought. Various parameterizations of the effect of sildenafil, sildenafil’s metabolite, and combinations of the two were examined. The parameterization identified as giving the largest increase in goodness of fit was a nonlinear function utilizing only the plasma concentration of sildenafil’s active metabolite [SldM]:

$$TVCL = \theta_{CL} \left(1 - \frac{[SldM]}{\theta_{Sld} + [SldM]}\right)$$

(10)

Final parameter values for this model are listed in Table IV. Bootstrapping with replacement demonstrated a good model fit, mean parameter values remained near the final model, and CV% was low. A posterior predictive check also demonstrated a good model fit. Fig. 5 demonstrates that most of the observed data points fall within the 90% prediction interval.

The resulting box-plot summaries for the PK-PD simulation prediction errors on clearance and EC50 are illustrated in Figs. 6 and 7, respectively. The first four sampling scenarios with post-PD sampling (6 h or later) demonstrated that improved predictability of clearance occurred when more post-PD samples were obtained. Notably, improvements in both bias and precision were seen as the number of samples increased. However, sampling scenario 4 with eight PK samples was only slightly better than scenario 3 with only three post-PD samples. Sampling on another occasion, even with eight samples spaced throughout the experiment at equivalent times to scenario 5, was no better than obtaining no samples at all (comparable bias and precision for scenarios 5 and 8). Scenario 6 represented the hypothetical best-case situation if dense blood samples were taken throughout the experiment. The bias was negligible and the precision was the tightest for this scenario; however, the bias was comparable and the precision only slightly better than both scenarios 3 and 4. Adding a single sample during the PD portion of the experiment (0–6 h) in scenario 7 was of no added value in comparison with scenario 2 where two samples were taken after the PD experiment. Finally, scenario 8 served as a point of comparison where no PK samples were taken. As expected, this produced the most imprecise estimates and introduced the largest amount of prediction error, along with scenario 5 where blood samples were taken in the same animal, but on a separate occasion. There was no bias in this schedule as the typical parameter values were assumed for all individuals. PD simulations reflected these conclusions for the three simulated sampling schedules.

**DISCUSSION**

The development of a PK model identifying a PK interaction between pregabalin and sildenafil in rats was investigated. Furthermore, these results served to inform the design of a future quantitative PD study examining the interaction of these drugs. This PK study design also serves as the first step toward identifying and quantifying a synergistic PD interaction between pregabalin and sildenafil.

Fundamental to identifying the source for an observed interaction between two compounds is to take the influence of a potential PK interaction into account (14, 15). The results of this analysis served that purpose by quantifying the PK of pregabalin with and without the presence of sildenafil in rats, noting a 30.2% decrease in the clearance of pregabalin when co-administered with sildenafil. This decrease in clearance was unexpected, and the mechanism for the PK interaction remains unclear. One potential hypothesis to explain this interaction may be an effect of sildenafil on renal transporters affecting pregabalin, although there is no information to enlighten this theory. The decrease in pregabalin clearance did produce a slight increase in the observed plasma concentrations of pregabalin over the observation period of interest.

The effect of pregabalin on the PK of sildenafil was not investigated. N-methyl sildenafil attained a $C_{\text{max}}$ of 2,100 ng/mL at 4–7 h, compared to sildenafil with a $C_{\text{max}}$ of 2,400 ng/mL at 2–4 h. While the metabolite $C_{\text{max}}$ was slightly lower

![Fig. 4. Scatter plots demonstrating the effect of sildenafil administration on clearance: A pregabalin clearance vs. a binary variable representing either no sildenafil (0) or with sildenafil (1) and B pregabalin clearance vs. the concentration of sildenafil’s metabolite.](image-url)
Fig. 7. Box plot summaries of prediction errors on EC50 for three of the eight sampling schedule scenarios: 3. (6, 8, and 24 h), 6. (1, 2, 4, 6, 8, 12, 18, and 24 h), and 8. no PK samples.
than sildenafil, the ratio between metabolite/parent (0.90) was much higher than that observed for the metabolite and parent in humans (0.4) (48). While the higher ratio may influence model development and subsequent human scale-up for non-saturating doses of sildenafil, the bolus and maintenance dose selected within this study attained a saturable plasma concentration of sildenafil for the PDE-5 receptor (≥3.5 nM) at all time points (49). As concentrations of sildenafil that would saturate the PDE-5 enzyme were attained, any minor impact of pregabalin on the PK of sildenafil would be negligible in this study design.

An additional objective was to determine a PK sampling strategy sufficient for representing actual rat PK profiles during future planned PD experiments. Capturing individual PK would not be straight-forward, as blood sampling during the PD experiment could lead to an introduction of PD measurement errors. For this reason, simulations provide an ideal platform for evaluating multiple limited sampling strategies for describing individual PK (17, 44, 50). Furthermore, this type of preclinical trial simulation allows efficient examinations of multiple experimental scenarios before any animals are utilized, ensuring that experimental animals are utilized in the most effective manner possible.

A total of eight sampling scenarios were examined by combining PK experimental data with simulated rat pregabalin plasma concentrations based on the above developed pregabalin PK model. Sampling strategy success was then assessed by comparing the PE of the predicted clearance to the actual (combined experimental and simulated data set) population clearance. Sampling the same rats on another occasion was determined to be of no added value compared to taking no blood samples at all due to the large degree of intra-individual and inter-occasion variability identified in this PK model for the volume (0.015(9.0%)) and clearance (0.098 (27.1)%). PK predictions using no blood samples from any of the rats produced no bias in the estimates of clearance, but it introduced a lack of precision in estimating individual PK. Optimally, taking eight blood samples spaced during and after the experiment produced the lowest bias and highest precision in the prediction errors.

Sampling in this optimal fashion described above served only as a benchmark, since collecting blood samples during the PD experiment would confound the PD experimental results. For this reason, additional sampling strategies involving a varying number of post-PD samples were also examined. Collecting one post-PD blood sample offered a small improvement in precision over no blood sampling, but it also introduced bias into the prediction errors. Increasing the number of post-PD samples reduced this bias and improved the precision; however, the degree of improvement in both bias and precision diminished with increasing blood samples. Increasing from three to eight post-PD samples slightly improved prediction error precision, but it was decided that these improvements in precision were not sufficient to warrant the collection of five additional samples due to the resulting blood volume alterations, animal stress, and research costs. Thus, collecting three post-PD blood samples was deemed sufficient for removing bias in the prediction error and drastically improving the precision.

Static allodynia PD simulations supported this conclusion by demonstrating that the three-sample schedule provided increased precision in estimating EC50s compared to taking no blood samples, and was only marginally worse than the unrealistic best-case scenario of eight samples taken during and after the PD experiment. This increased precision in identifying individual EC50s will allow a greater ability to identify changes in potency when examining the interaction between these two compounds. Furthermore, these types of simulations may also serve as an example for other preclinical studies to help refine experimental design.

Overall, this study developed a two-compartment model capable of accurately describing pregabalin PK in the presence or absence of sildenafil. These results also served to inform the sampling schedule design for future quantitative PK-PD study examining the interaction of these two drugs. Finally, this PK study also served as an initial step towards identifying and quantifying potential PK and PD interactions between pregabalin and sildenafil.

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