Behavioral evidence for pattern separation in human episodic memory

Ewa Zotow, James A Bisby, Neil Burgess

Institute of Cognitive Neuroscience, University College London, London WC1N 3AZ, United Kingdom; Institute of Neurology, University College London, London WC1N 3BG, United Kingdom; Division of Psychiatry, University College London, London W1T 7BN, United Kingdom

An essential feature of episodic memory is the ability to recall the multiple elements relating to one event from the multitude of elements relating to other, potentially similar events. Hippocampal pattern separation is thought to play a fundamental role in this process, by orthogonalizing the representations of overlapping events during encoding, to reduce interference between them during the process of pattern completion by which one or other is recalled. We introduce a new paradigm to test the hypothesis that similar memories, but not unrelated memories, are actively separated at encoding. Participants memorized events which were either unique or shared a common element with another event (paired “overlapping” events). We used a measure of dependency, originally devised to measure pattern completion, to quantify how much the probability of successfully retrieving associations from one event depends on successful retrieval of associations from the same event, an unrelated event or the overlapping event. In two experiments, we saw that within event retrievals were highly dependent, indicating pattern completion; retrievals from unrelated events were independent; and retrievals from overlapping events were antidependent (i.e., less than independent), indicating pattern separation. This suggests that representations of similar (overlapping) memories are actively separated, resulting in lowered dependency of retrieval performance between them, as would be predicted by the pattern separation account.
environment (Kyle et al. 2015). The hippocampal activity patterns associated with overlapping routes (Chanales et al. 2017) and narratives with overlapping elements (Chadwick et al. 2011; Mliivojevic et al. 2016) were found to diverge over time. However, these tasks either did not provide a behavioral read-out or did not look at memory for multiple event elements.

In the current study, we focused on the role of pattern separation in preventing incorrect pattern completion in episodic memory. While the elements of each event should be bound together into coherent representations, resulting in a holistic retrieval of each event, the patterns associated with different events should be sufficiently distinct to prevent interference. Recent studies have shown that multielement events (e.g., person—location—object) are stored and retrieved as single coherent representations. That is, successful retrieval of one association from an event is statistically related to retrieval of all other associations from the same event (Horner and Burgess 2013, 2014; Horner et al. 2015; Bisby et al. 2018). This within-event dependency across retrievals and the holistic manner in which events are reexperienced is consistent with pattern completion. In addition, neuroimaging evidence has shown that the individual elements from an event are represented by activity in separate neocortical regions; at retrieval all elements, including the non-cue/non-target item, are reinstated when a single association is tested, and this reinstatement is supported by the hippocampus (Horner et al. 2015).

The observed within-event dependency across retrievals of associations from events offers a potential basis to examine pattern separation processes at the behavioral level. For instance, while retrievals of associations within an event are related, retrieval success between separate unrelated events should be independent. Further, as pattern separation aims to reduce interference of overlapping memories by making their representations less similar, the retrieval success across two events that share a common element (e.g., the same object used in two separate events) should show a decreased dependency if subject to pattern separation at encoding. That is, the probability of making a successful retrieval from one event should be dependent on the retrieval of all other elements from that event but independent from retrieval success of elements from an unrelated event, and “negatively dependent” (i.e., showing decreased dependency) from a related event with which it shares an element.

Across two experiments, we aimed to assess the pattern of retrievals both within and across events for evidence of pattern separation- and pattern completion-like processes. Similar to previous work (Horner and Burgess 2013, 2014; Horner et al. 2015; Bisby et al. 2018; Joensen et al. 2019; Ngo et al. 2019), we instructed participants to encode multielement events involving a person, a location and an object. Some of the events overlapped with another event, that is, they shared a common element (e.g., the same person was part of both events). Other events were fully unique (nonoverlapping) and didn’t share any elements with other events (Fig. 1). At test, associative accuracy was assessed using a six-alternative forced choice test.

We assessed the statistical dependency of success in retrieving different associations from either the same event or from different events, where within-event dependency is thought to reflect pattern completion—a common process in retrieving, for example, both the location and the person in an event, when cued by the object (see Horner and Burgess 2013, 2014). Here we are additionally interested in the dependency between retrievals from different events as a function of whether or not they share an element. For instance, the probability of retrieving a person when cued by an object for both event 1 and event 2 (where event 1 and event 2 could be either unrelated or overlapping). The dependency measure reflected how often these associations are retrieved either both correctly or both incorrectly.

To account for the performance levels of each participant across different types of element, and levels of guessing, the dependency measure in each analysis was compared to independent and dependent models, predicting respectively the level of dependency expected from independent pairwise associations and from associations modulated by a common factor for each event (estimated from other within-event associations; Horner and Burgess 2014).

According to the pattern separation account of memory formation, the representations of overlapping events will actively separate to prevent interference, leading to reduced dependency in retrievals from one event on retrievals from the other event than that observed for unrelated, nonoverlapping events.

Results

Experiment 1

Associative accuracy

Overall accuracy across all trials was good (66.42%, SD = 17.66) and well above chance (chance would be 16.7% given the six test options). Analysis of performance across cue-type (collapsed across retrieval-type) using a $2 \times 3$ repeated-measures ANOVA with factors of event type (overlapping, nonoverlapping) and cue-type (person, location, object) showed a significant main effect of event type ($F_{(2,58)} = 13.41, P = 0.001, \eta^2 = 0.32$) with slightly higher accuracy in overlapping events. There was no significant main effect of cue type ($F_{(2,58)} = 2.56, P = 0.09, \eta^2 = 0.08$) nor a cue-type × event type interaction ($F_{(2,58)} = 0.07, P = 0.93, \eta^2 < 0.01$). A similar $2 \times 3$
ANOVA on target-type (collapsed across cue-type) showed no main effect of retrieval type \((F_{(2,58)} = 2.31, P = 0.11, \eta^2 = 0.07)\). The main effect of event was the same as in the cue-type analysis. There was no interaction between retrieved type and event type \((F_{(2,58)} = 1.96, P = 0.15, \eta^2 = 0.06)\); (see Table 1).

**Within-events dependency**

Dependency was assessed by constructing contingency tables for retrieving two elements when cued with the third element, and retrieving one element when cued by the other two elements across retrieval trials (see Materials and Methods). We then calculated within-event dependency \((D)\) in the data by taking the proportion of events where elements were both correctly or incorrectly retrieved (Table 3). This dependency was then compared to the amount of dependency predicted if retrievals from the same event were completely independent \((D_{Di})\) or dependent \((D_{Dd})\; (\text{see Materials and Methods for more information on how the models were constructed}). We compared dependency in the data with both independent and dependent models separately for nonoverlapping and overlapping events. A \(2 \times 3\) repeated-measures ANOVA with factors of event type (nonoverlapping, overlapping) and dependency measure \((D, D_{Di}, D_{Dd})\) showed a significant interaction between event type and dependency \((F_{(1.52,9.82)} = 9.82, P = 0.001, \eta^2 = 0.25;\) Greenhouse-Geisser corrected).

To further assess the interaction, we first analyzed overlapping and nonoverlapping events separately, comparing the data to independent and dependent models. For nonoverlapping events, we found greater within-event dependency in the data compared to the independent model \((D > D_{Di}, t_{(29)} = 7.41, P < 0.001, d = 1.33)\) and less than the dependent model \((D < D_{Dd}, t_{(29)} = 4.06, P < 0.001, d = 0.75)\). Similarly for overlapping events, dependency in the data was greater than the independent model \((D > D_{Di}, t_{(29)} = 9.00, P < 0.001, d = 1.66)\) and was less than the dependent model \((D < D_{Dd}, t_{(29)} = 10.46, P < 0.001, d = 1.67)\) (Fig. 2).

We next performed a direct comparison of dependency between nonoverlapping and overlapping events. We calculated a difference score between each participant’s dependency and the independent model \((D_{Di})\) and between each participant’s dependency and the dependent model \((D_{Dd})\). This reflected a relative dependency level; positive scores indicated a higher dependency level than predicted by the average performance under the relevant model. We then compared the relative dependency scores between nonoverlapping and overlapping events.

There was no difference between nonoverlapping and overlapping events in the amount of dependency relative to the independent model \((D_{Di}, t_{(29)} = 1.87, P = 0.077, d = 0.34)\). A comparison of dependency relative to the dependent model showed a significant difference between events \((D_{Dd}, t_{(29)} = 3.04, P = 0.005, d = 0.55)\) with less dependency in the data compared to the dependent model for overlapping events. Importantly, both nonoverlapping and overlapping events showed greater dependency in the data than predicted by the independent model.

### Table 1. Experiment 1: Proportion correct (SD) for associative memory performance across nonoverlapping and overlapping events for each cue and retrieval type

| Cue type | Retrieved type | Location | Retrieved type | Location |
|----------|----------------|----------|----------------|----------|
|          | Nonoverlap | Overlap  | Nonoverlap | Overlap  |
| Person   | n/a         | n/a      | 0.63 (0.17) | 0.68 (0.21) |
| Location | 0.64 (0.19) | 0.68 (0.22) | n/a | n/a |
| Object   | 0.62 (0.18) | 0.72 (0.22) | 0.62 (0.20) | 0.64 (0.24) |

**Dependency in matched events**

As we have proposed, successful pattern separation between events that share a common element would likely reduce dependency due to the decorrelation of neural representations of those events. To examine this prediction, we examined the amount of dependency across matched overlapping events that shared a common element (i.e., whether retrieval success of an association from one event is dependent on the retrieval success of the matching association from an overlapping event) and compared this to the amount of dependency across nonoverlapping events. Associations where the common element was the cue were not included in this analysis.

We calculated dependency in the data and corresponding independent and dependent models across pairs of matched, overlapping events (see Materials and Methods for how this was achieved).

First, we looked at whether dependency in retrievals from overlapping events differed from nonoverlapping events. We ran a \(2 \times 3\) repeated-measures ANOVA with factors event type (nonoverlapping, overlapping) and dependency measure \((D, D_{Di}, D_{Dd})\). It is important to note that the measures used here are the within-event dependency measures for the nonoverlapping events and the across-events dependency measures for the overlapping events. As expected, we found a significant interaction \((F_{(1.46,42.24)} = 37.91, P < 0.001, \eta^2 = 0.57;\) Greenhouse-Geisser corrected).

To test the hypothesis that the overlap leads to a lowered dependency across matched events, we compared the amount of across-event dependency to its independent model for overlapping, matched events. We tested whether the retrieval of associations from one event depends on the retrieval of associations from its matched event. As predicted, this across-events dependency was lower than the independent model \((D < D_{Di}, t_{(29)} = 2.70, P = 0.011, d = 0.48),\) and lower than the dependent model \((D < D_{Dd}, t_{(29)} = 6.48, P < 0.001, d = 1.17)\) (third set of three bars in Fig. 3). This pattern is different to the one observed in the within-events dependency comparisons where the dependency in data was higher than predicted by the independent model (Fig. 2). The lower than baseline (independent model) dependency supports the hypothesis that the representations of similar events are actively separated.

Across participants, the level of relative dependency also was highly correlated with the interference measure, that is, the correlation between accuracy on overlapping pairs of events \((r = 0.80, P < 0.001)\).

Next, we directly compared the dependency between nonoverlapping and overlapping events. Here again, we used the within-event dependency measures for the nonoverlapping events and the across-events dependency measures for the overlapping events. The level of dependency relative to the independent model was lower for the overlapping events \((D_{Di}, t_{(29)} = 6.90, P < 0.001, d = 1.26)\). A comparison of dependency relative to the dependent model also showed a significant difference between events \((D_{Dd}, t_{(29)} = 4.66, P < 0.001, d = 0.85)\), with lower dependency as compared to the dependent model in the overlapping events.
Comparison of across-events relative dependencies

Crucially, we wanted to see whether the dependency score relative to the independent model for the matched pairs of events differed from the relative dependency for unrelated pairs. If matched events are stored more independently than what would be expected from unrelated memories, this would provide evidence for pattern separation of overlapping episodic memories. To test this, we randomly paired unrelated events (separately for events of nonoverlapping and overlapping type). Unrelated events of overlapping type were the events that did have a matched pair, but here were paired up with another (unrelated) event to calculate the across-events dependency. For example, if event1, overlapped with event2, in this analysis it could be randomly paired with any other event from the set of overlapping paired events (e.g., event3) but not with the event2. This created pairs of events that were of unrelated, overlapping type with no shared elements (in contrast to pairs of matched events which shared one common element with each other).

We then calculated the across-events dependency measures for these pairs in the same way as for the matched pairs of events (see above). This was to ensure that the lowered dependency is not a result of the event type, but rather that it is specific to pairs of matched events with a common element. In other words, we wanted to see whether the need for active separation of two related episodes leads to a “negative” relative dependency between them.

As each analysis used a randomly created set of unrelated event pairs, the specific pairs selected in any given analysis could have affected the results. We therefore ran each analysis 1000 times and used a Brown’s method (i.e., an extension to the Fisher’s combined probability test; Brown 1975; Poole et al. 2016) which provides a combined $P$-value for nonindependent tests.

Each analysis compared the relative dependency in the matched pairs (which was the same in all 1000 analyses, as these pairs were not randomly selected) to the relative dependency in the same number of unrelated pairs of events, with different unrelated pairs selected each time.

As would be predicted by a pattern separation account, the matched events showed lower relative dependency than unrelated events of nonoverlapping type, $P<0.008$, and of overlapping type, $P<0.006$ (Fig. 3). The two types of unrelated events (i.e., chosen from overlapping or nonoverlapping events) were not significantly different from each other, $P=0.532$, and neither was significantly different from their baseline, i.e., the independent model ($P=0.497$ for unrelated nonoverlapping events and $P=0.475$ for unrelated overlapping events; Fig. 3).

Experiment 2

Experiment 1 compared the dependency in retrieval of associations from events that were fully unique or that shared a common element with another event (i.e., overlapping events). The results were consistent with pattern separation of overlapping memories. Experiment 2 set out to replicate these behavioral results while measuring brain activity in an fMRI task (not reported here).

Figure 2. Mean dependency for the observed data, the independent model and the dependent model for the within-event analysis. The error bars represent ±1 SEM. (*) $P<0.001$.

Figure 3. Mean across-event dependency for pairs of unrelated events (first set of three bars: unrelat- ed, nonoverlapping events; second set of three bars: unrelated, overlapping events) and pairs of matched, overlapping events (third set of three bars). The data for unrelated pairs was obtained by randomly selecting pairs of unrelated events (separately for overlapping and nonoverlapping events) and calculating the dependency and the models, and repeating this procedure 1000 times, each time with a different set of random pairs. The $P$-values were obtained using Brown’s method for combining tests of significance (Brown 1975). The bars represent the mean relative dependency, and the error bars represent the mean standard error across the 1000 tests (unrelated events) or ±1 standard error (matched events). (*) $P<0.01$; (**) $P<0.001$.

Comparative accuracy

Overall accuracy was slightly lower than in Experiment 1, at 54.85% (SD = 19.43), but still well above the chance level of 16.7%. Analysis of performance across cue-types (collapsed across retrieval-types) using a one-way repeated-measures ANOVA showed no effect of cue type ($F_{(2,58)} = 0.03$, $P = 0.967$, $\eta^2 = 0.001$), and analysis of performance across retrieved-type (collapsed across cue-type) also showed no effect of retrieved-type ($F_{(2,58)} = 1.60$, $P = 0.212$, $\eta^2 = 0.09$). The means and standard deviations for accuracy across different cue and retrieval types for both conditions are in Table 2.

Within-events dependency

As in Experiment 1, the within-events dependency was assessed by constructing contingency tables for retrieving two elements when cued with the third element, and retrieving one element when cued by the other two elements across retrieval trials (see Materials and Methods of Experiment 1). The within-event dependency ($D$) in the data was calculated as the proportion of events where elements were both correctly or incorrectly retrieved. This dependency measure was then compared to the amount of dependency predicted if retrievals from the same event were completely independent ($D_i$) or dependent ($D_d$; see Materials...
and Methods of Experiment 1 for more information on how the models were constructed).

The within-event dependency in the data was greater than the independent model (D > Dd, t(29) = 7.38, P < 0.001, d = 1.35) and lower than the dependent model (D < Dd, t(29) = 11.09, P < 0.001, d = 2.02) (Fig. 4).

**Dependency in matched events**

As in Experiment 1, we examined the amount of dependency across matched overlapping events that shared a common element (i.e., whether retrieval success of an element from one event is dependent on the retrieval success of an element from an overlapping event). We calculated dependency in the data and corresponding independent and dependent models across pairs of events (see Materials and Methods of Experiment 1).

Consistently with the behavioral study, the across-events dependency was lower than the independent model (D < Dd, t(29) = 3.04, P = 0.001, d = 0.56), and lower than the dependent model (D < Dd, t(29) = 8.54, P < 0.001, d = 1.56) (Fig. 5).

**Comparison of across-events relative dependencies**

As in Experiment 1, we calculated a relative dependency measure (difference score between each participant’s dependency and its independent model: D-Di). We looked at whether the relative dependency score is lower for overlapping events than for pairs of unrelated events. As before, in order to create unrelated pairs, we randomly selected sets of two events with no shared elements. We then calculated the across-events dependency measure for these pairs in the same way as for the matched pairs of events (see Materials and Methods of Experiment 1). Because each analysis selected a random set of unrelated pairs, we repeated the procedure 1000 times, with different unrelated pairs selected each time, and used the Brown’s method to combine the probability values (Brown 1975; Poole et al. 2016).

Supporting the results from Experiment 1, we found that overlapping events had lower relative dependency than the unrelated events (P = 0.001). The dependency in unrelated events was not different from baseline, i.e., the independent model (P = 0.289) (Fig. 5).

**Discussion**

Computational models of hippocampal function propose that similar memories are stored as distinct nonoverlapping representations through a process of pattern separation (Marr 1971). This process is complementary to pattern completion by which a subset of cues from a previous experience can reactivate the stored pattern representing that experience (Marr 1971; Hopfield 1982; McClelland et al. 1995; Rolls 2015). These processes therefore give rise to two complementary characteristics of memory such that even similar events are stored as independent nonoverlapping representations, and that all elements of an event can be retrieved in a holistic manner. To test these hypotheses, we examined whether events that share a common element would be represented as more distinct than unrelated events, as assessed by the relationships between retrievals from within and between the overlapping or unrelated events.

Since the role of pattern separation at encoding is to reduce interference during pattern completion at retrieval, we adapted a pattern completion task to provide a measure of both processes. Two experiments provided consistent results. While retrievals from within the same event show increased statistical dependency (as predicted by pattern completion), retrievals from overlapping events showed lower dependency than expected from unrelated events. This supports the proposal that pattern separation serves to decorrelate similar inputs by transforming their representations into orthogonal patterns (Marr 1971; McClelland et al. 1995).

Our results are consistent with the view that events are stored as coherent representations and retrieved in a holistic manner. In accordance with previous studies, we found greater dependency in the retrieval of elements from the same event (Horner and Burgess 2013, 2014; Horner et al. 2015; Bisby et al. 2018), supporting the holistic way in which episodic memories are stored and retrieved (Figs. 2, 4; Tulving 1983). We assume that within-event dependency reflects the associative structure in which event elements are bound together into single representation. When cued by a single event element, all associated within-event elements are reinstated. This is consistent with computational models of hippocampal pattern completion and the way event elements are stored in an autoassociative network in which presentation of a partial cue will cause reinstatement of all associated event elements (Marr 1971; Hopfield 1982; McClelland et al. 1995). Neuroimaging evidence, using a similar task and dependency measure as we used here, is also complementary to our results in suggesting that all within-event elements are reinstated in neocortical areas and this reinstatement is supported by the hippocampus (Horner et al. 2015) and more specifically, region CA3 (Grande et al. 2019).

Within-event dependency was also seen in events even when they overlapped with other events (i.e., they shared a common element). This finding is important as it suggests that, while pattern separation should specifically affect representations that may interfere with each other due to their similarity (i.e., representations of overlapping events), the constituent elements of each of these events, which needs to be bound into coherent narratives, are not separated but instead continue to show pattern completion. This is consistent with findings from rats which showed that the hippocampus binds information encountered in the same context together while separating events from different contexts (Wills et al. 2005; McKenzie et al. 2014).

While retrieval of within-event elements should show high dependency, we should not observe the same pattern of results
in across-event dependency (i.e., the extent to which retrieving an association from one event depends on the retrieval of associations from another event). If events are represented as separate engrams, we would expect that the dependency in the retrieval of their respective associations will be independent, that is, as predicted by the independent model. For unrelated pairs of events this was indeed found to be the case (Figs. 3, 5).

Interestingly, and crucially for our hypothesis, retrievals across overlapping events (i.e., events with one element in common) showed a level of dependency that was significantly lower than predicted by the independent model (Figs. 3, 5). For dependency to be less than expected by the independent model, the probability of retrieving an association from one event must be negatively related to the probability of retrieving an association from its overlapping event. Although overlapping events shared an element and so might be expected to become associated during learning, the success on the task depended on the ability to discriminate between the two events in order to avoid interference, which we propose is achieved by an increased separation of their respective representations. The finding that the dependency was “negative” only for the matched pairs from overlapping events, but not for pairs from unrelated overlapping events, suggests that the results are not simply due to the fact that events that are not fully unique are processed in a different way; rather, the increased separation occurs only for the specific pairs of events which share an element in common and not for all events that happen to share a common element with another event. The role of pattern separation may be therefore to accentuate the differences between similar inputs to avoid subsequent interference. This is performed by the dentate gyrus which decorrelates the representations of similar memories by “selecting” a different population of CA3 neurons for their storage, driving their respective representations further apart (Marr 1971; Treves et al. 2008). As a result, similar memories have less similar representations than unrelated memories do on average. These two mechanisms ensure that the overlapping representations are less likely to interfere with each other while their constituent elements are bound together into single representations. This shows the complementary functions of pattern separation and completion in supporting episodic memory.

Our paradigm extends upon research examining pattern separation processes in humans by using a task that segregates the encoding and retrieval phases, and distinguish between successful encoding and pattern separation. Pattern separation and pattern completion operate at, respectively, encoding and retrieval, and so it has been suggested that an appropriate task should be one where the process of interest (separation or completion) is the most appropriate strategy (Hunsaker and Kesner 2013; Liu et al. 2016). The continuous version of the MST requires “recall-to-reject” in addition to intentional encoding, making it difficult to separate the two processes, an issue recognized by the authors (Kirwan and Stark 2007). The study-test version, although separates encoding and retrieval, still does not examine whether the encoded representations of similar items are more separated than those of dissimilar items, and may instead reflect the accuracy of encoding of the previously seen items.

Our findings are consistent with a study where similar scenes were paired with either two different facial stimuli, and so needed to be discriminated, or with the same face. The demand to distinguish between the similar scenes led to reduced interference of those scenes in a subsequent association task with novel stimuli (Favila et al. 2016). The need for discrimination and the subsequent reduction in interference were both related to decreased similarity in the scenes’ hippocampal representations assessed by fMRI. This is consistent with the view that pattern separation serves to reduce interference, in the context of separation of similar static scenes rather than the multielement events studied here. The current task provides another way of looking at behavioral pattern separation while overcoming some of the limitations of pattern separation/completion tasks outlined by Liu et al. (2016) and others (Kirwan and Stark 2007; Hunsaker and Kesner 2013).

The current task targets the associative role of the hippocampus in multielement episodic memory. The hippocampus is
thought to be specifically required for contextual or relational memory as opposed to memory for individual items or a feeling of familiarity (Uncapher and Rugg 2005; Horner et al. 2012; Aggleton and Brown 1999). Item memory, on the other hand, may be processed in other medial temporal lobe structures (e.g., perirhinal cortex) (McClelland et al. 1995; Eichenbaum et al. 2007), but see also (Squire and Zola-Morgan 1991). Although the use of an old/new recognition paradigm with similar foils may require the hippocampus for recollection of small details rather than familiarity (Holdstock et al. 2002), it does not require pattern separation at encoding and so does not directly address the formation of separated representations for storage to avoid interference during recall.

In conclusion, we present a novel and relatively process-pure way of behaviorally investigating pattern separation. We propose that the lowered (“negative”) dependency in retrieval of associations from overlapping events results from the need to differentiate their neural representations through pattern separation. This finding is specific to pairs of events with a common element and not to pairs of unrelated events (whose retrieval was found to be independent from each other), which confirms the predictions regarding how similar episodic memories are represented.

Materials and Methods

Experiment 1

Participants

Thirty-nine participants were recruited from the University College London student population. Seven participants were excluded due to poor overall task accuracy (below 25%) and two due to too high accuracy (above 90%); very high performance resulted in too low variability to give reliable results in the dependency analysis (see below). The remaining 30 participants (25 females, 5 males; sample size exceeds related behavioral studies, Horner and Burgess 2013, 2014) had a mean age of 25.61 (SD = 6.22), had normal or corrected-to-normal vision, were fluent English speakers and were familiar with Western culture including major celebrities and politicians (self-reported). The study was approved by the UCL Research Ethics Committee and all participants provided written informed consent prior to taking part in the study. Following completion of the test, participants were debriefed and paid for their time.

Materials

Word stimuli included 40 famous people (e.g., Tom Cruise, Barack Obama), 40 locations (e.g., supermarket, kitchen) and 40 objects (e.g., necklace, bottle). Before attending the study, participants selected famous people they were familiar with from a list of 60 candidates; out of all familiar to the participant, 40 were chosen at random (all participants were familiar with at least 40 people from the full list). Stimuli were combined to create 45 three-element events with each event consisting of a person, location, and object. Fifteen of these events were unique from each other in that they shared no common elements (i.e., nonoverlapping events; Event 3 in Fig. 1A). The remaining 30 events were combined to create 15 sets of two events in which the two events of a set shared one common element (i.e., overlapping events; Events 1 and 2 in Fig. 1A). Pairs of overlapping events with a shared element will be referred to as matched events. For matched events, the common element (person, location, or object) was counterbalanced across event sets. Novel randomized events were created for each participant. The Cogent 2000 toolbox (www.fil.ion.ucl.ac.uk) for Matlab R2016a (MathWorks) was used for stimulus presentation and data collection. We used white text on gray background, and Helvetica font of size 30.

Procedure

The experiment consisted of a study phase and a test phase (Fig. 1B, C). During the study phase, participants encoded a total of 45 event triplets with each comprising a person, location and object. Each trial began with a 0.5 sec fixation period, after which participants were presented with one of the event triplets (Fig. 1A). All three elements remained on the screen for 8 sec and participants were instructed to imagine all event elements interacting as vividly as possible. The screen location of each element type (person, location, object) was randomized across encoding trials. The trial ended with a blank screen presented for 1.5 sec. Events were only shown once during the study phase.

At test, associative accuracy was assessed using a six alternative forced choice test. On each trial, following a 0.5 sec fixation, participants were presented with one of the previously encoded items at the top of the screen (cue) and six possible “target” items were presented below (Fig. 1B). They were then instructed to decide which of the targets had previously appeared in the same event as the cue. All targets on a single test trial were of the same category and all appeared in the study phase as elements of different events. For example, if cueing with the person to retrieve the associated location, all six options were of locations. Participants were given 8 sec to select their option via button press. All possible cue-target pairs were tested in both directions (e.g., cue with the person to retrieve the location, cue with the location to retrieve the person), giving a total of 270 trials. The order of trials was randomized. Each trial terminated with a 1.5 sec blank screen.

Analysis of within-event dependency

We created 2 x 2 contingency tables for each participant. For within-event dependency, the contingency tables were based on (i) the probability of retrieving two items from the same event when cued by the remaining item from that event (ABAc; e.g., retrieving either a person “b” or a location “c” when cued by an object “A”), and (ii) the probability of retrieving an item when cued by the two remaining items from the event (BaCa; e.g., retrieving a person “a” when cued either by a location “B” or an object “C”).

This measure was then compared to an independent model and a dependent model calculated individually for each participant. This within-subject comparison accounted for individual differences in performance on the raw dependency score. The models estimated the level of dependency based on the average performance and level of guessing. The independent model assumed that the retrieval of any two elements from the same event is completely independent and is contingent only on the overall accuracy level. It was calculated by multiplying the probabilities of separately retrieving two elements from the categories in question.

The dependent model additionally adjusted the predicted level of dependency by an event-specific “episodic factor”—a measure of the average performance on a given event (across all other retrieval trials for that event) relative to the overall performance. The probability of retrieving any association from an event was weighted by the episodic factor for that event. The dependent model also accounted for the level of guessing: the episodic factor affected the probability of intentional retrieval but not the probability of correct guessing, which is assumed to be independent (see Table 3).

For each participant, we calculated the measure of dependency as well as the independent and dependent models separately for events of nonoverlapping and overlapping types. This was to check whether both types are retrieved as complete events (i.e., both show high within-event dependency).

Analysis of across-event dependency

The contingency tables for dependencies across pairs of matched events were based on the probability of correctly or incorrectly retrieving associations from both events. Unlike in the “standard” within-event dependency described above, where both compared associations (e.g., Ab and Ac) came from the same event, each association came from a different event from a pair (event1 and event2). Importantly, for calculation of the dependency measure and
Table 3. Contingency tables for independent and dependent models giving the frequency (over events) of the four combinations of correct or incorrect retrievals of elements B and C when cued with element A

| Retrieval of element (C) | Correct (P_{AB}) | Incorrect (1 - P_{AB}) |
|--------------------------|------------------|-----------------------|
| Independent model        |                  |                       |
| Correct (P_{AC})         | $\sum_{i=1}^{N} p_{AB}p_{AC}$ | $\sum_{i=1}^{N} p_{AB}(1 - p_{AC})$ |
| Incorrect (1 - P_{AC})   | $\sum_{i=1}^{N} p_{AB}(1 - p_{AC})$ | $\sum_{i=1}^{N} (1 - p_{AB})$ |
| Dependent model          |                  |                       |
| Correct (P_{AC})         | $\sum_{i=1}^{N} \tilde{p}_{AB}p_{AC}$ | $\sum_{i=1}^{N} \tilde{p}_{AB}(1 - p_{AC})$ |
| Incorrect (1 - P_{AC})   | $\sum_{i=1}^{N} \tilde{p}_{AB}(1 - p_{AC})$ | $\sum_{i=1}^{N} (1 - \tilde{p}_{AB})$ |

Dependent model replaces the probability of correctly recalling B when cued with A (across all events; $p_{AB}$) with $\tilde{p}_{AB} = E_{i}(p_{AB} - P_{c}/c) + P_{c}/c$ where the episodic factor $E_{i}$ reflects performance on event i relative to other events (based on retrievals other than B and C cued by A), $P_{c}$ is the probability of guessing, and $c = 6$ is the number of choices in the test trial. The level of guessing is unknown but we can assume that all errors reflect guessing (Horner and Burgess 2013), and so $P_{c} = (1 - (p_{AB} + p_{BA} + p_{BC} + P_{AC} + P_{CA})/6)/c(c - 1)$. $P_{AC}$ is replaced similarly. The dependency model equates to the independent model if the episodic factors are set to 1.

models we excluded the trials in which the overlapping element was a cue, as these types of trials had two possible correct answers, and potentially a retrieval of the incorrect pair could interfere with the retrieval of the appropriate element. Here the contingency tables were based on (i) how the probability of retrieving a specific association depends on the probability of retrieving the same type of association from the corresponding paired event ($A_{B}B_{A2}$; where A is not the shared item), (ii) how the retrieval of an item from one event depends on the retrieval of a different-type item from the corresponding event when cued by the same-type items ($A_{B}A_{C2}$; where A is not the shared item), and (iii) the probability of retrieval of the same-type items when cued by different-type items in each event ($B_{A1}C_{A2}$; where A is the shared item, because pairs where the shared element is the cue were excluded).

Experiment 2

Participants

Thirty-three neurologically healthy participants were recruited from the student population at the University College London. Three participants were excluded due to low performance (<20% accuracy). The accuracy threshold was lowered to reflect the overall lower performance in this study, potentially caused by the different setup (sitting at a desk versus lying in an fMRI scanner). The remaining 30 participants (10 males) had a mean age 24.43 (SD = 3.87), normal or corrected-to-normal vision and were right handed. Participants were paid £20 for the scanning session. The study was approved by the University College London Research Ethics Committee (1825/003) and informed consent was obtained before the session. Following completion of the test, participants were debriefed and paid for their time.

Materials

The same stimuli as in Experiment 1 were used. As shown above, the dependency results for pairs of unrelated events were the same regardless of whether those pairs consisted of fully unique events or events which shared a common element with another event. Therefore, in this experiment only events of overlapping type were used, allowing for the use of a higher number of overlapping pairs while minimizing the overall time in the scanner. Word stimuli included 30 famous people (e.g., Tom Cruise, Barack Obama), 30 locations (e.g., supermarket, kitchen), and 30 objects (e.g., necklace, bottle). Participants selected famous people they were familiar with from a list of 60 candidates before attending the session; out of all familiar to the participant people, 30 were chosen at random (all participants were familiar with at least 30 people from the full list). Stimuli were combined to create 36 three-element events, giving 18 pairs with an overlapping element. The type of the common element was counterbalanced across event sets. Novel randomized events were created for each participant.

Procedure

Procedure was similar to Experiment 1, however as a result of this task being performed in an MRI scanner, the following aspects were changed: participants lay in a supine position with the visual display reflected from a back projector by a mirror attached to the head coil. During the test phase, participants responded using three keys with their right hand and three keys with their left hand.

The main part of the procedure was the same as in Experiment 1 and consisted of a study phase and a test phase. The study phase was the same as in Experiment 1 except it consisted of 36 trials. The test phase was also the same except it was divided into two scanning blocks with 108 trials in each. Pairwise associations for each event were split so that half were tested in the first run and half in the second run. The order within each run was randomized.

Analysis

All analyses of behavioral data were the same as described in Experiment 1, with the exception that here only events of overlapping type were used.

Acknowledgments

This work was supported by Wellcome and European Research Council advanced grant EUROMEM and a PhD studentship from the UCL Division of Psychology & Language Sciences. The Matlab script used to calculate the measure of dependency in data, independent and dependent models is provided by Dr Aidan Horner at https://osf.io/k495x/. The data sets and further information about the analysis are available on FigShare at https://doi.org/10.6084/m9.figshare.12407093.v1

REFERENCES

Aggleton J, Brown MW. 1999. Episodic memory, amnesia, and the hippocampal, anterior thalamic axis. Behav Brain Sci 22: 425–444. doi:10.1017/S0140525X99002034

Amit DJ, Gutfreund H, Sompolinsky H. 1987. Information storage in neural networks with low levels of activity. Phys Rev A 35: 2293. doi:10.1103/PhysRevA.35.2293

Bakker A, Kirwan BC, Miller M, Stark CEL. 2008. Pattern separation in the human hippocampal CA3 and dentate gyrus. Science 319: 1640–1642. doi:10.1126/science.1152862

Bisby JA, Horner AJ, Bush D, Burgess N. 2018. Negative emotional content disrupts the coherence of episodic memories. J Exp Psychol Gen 147: 243–256. doi:10.1037/xge0000356

Bonnici HM, Kumaran D, Chadwick MJ, Weiskopf N, Hassabis D, Maguire EA. 2012. Decoding representations of scenes in the medial temporal lobes. Hippocampus 22: 1143–1153. doi:10.1002/hipo.20960
Brown MB. 1975. A method for combining non-independent, one-sided tests of significance. *Biometrics* **31**: 987–992. doi:10.2307/25329826

Chadwick MJ, Hassabis D, Maguire EA. 2011. Decoding overlapping memories in the medial temporal lobes using high-resolution fMRI. *Learn Mem* **8**: 742–746. doi:10.1101/lm.256711

Chanales AH, Oza A, Pavlia SE, Kuhl BA. 2017. Overlap among spatial memories triggers divergence of hippocampal representations. *Curr Biol* **27**: 1–47. doi:10.1016/j.cub.2016.09.012

Cohen NJ, Eichenbaum H. 1993. Memory, amnesia, and the hippocampal system. *Clin Neurobiol Hippocampus An Interview* **55173**: 10.1093/acprof:oso/9780195923888.003.0003

Dasario AM. 1989. The brain binds entities and events by multiregional activation from convergence zones. *Neural Comput* **1**: 123–132. doi:10.1162/neco.1989.1.1.123

Davachi L. 2006. Item, context and relational episodic encoding in humans. *Curr Opin Neurobiol* **16**: 693–700. doi:10.1016/j.conb.2006.10.012

Eichenbaum H, Tonlewis AP, Ranganath C. 2007. The medial temporal lobe and recognition memory. *Annu Rev Neurosci* **25**: 123–152. doi:10.1146/annurev.neuro.30.051606.094328

Favila SE, Chanales AH, Kuhl BA. 2016. Experience-dependent hippocampal pattern separation prevents interference during subsequent learning. *Nat Commun* **7**: 1–10. doi:10.1038/ncomms11066

Gold AE, Kesner RP. 2005. The role of the CA3 subregion of the dorsal hippocampus in spatial pattern completion in the rat. *Hippocampus* **15**: 808–814. doi:10.1002/hipo.20103

Grande X, Berron D, Horner AJ, Bisty JA, Duzel E, Burgess N. 2019. Holistic recollection via pattern completion involves hippocampal subfield CA3. *J Neurosci* **39**: 8100–8111. doi:10.1523/JNEUROSCI.0722-19.2019

Holdstock JS, Mayes AR, Roberts N, Cezayirli E, Isaac CL, Reilly RCO, Kyle CT, Stokes JD, Lieberman JS, Hassan AS, Ekstrom AD. 2015. Successful retrieval of competing spatial environments in humans - a systematic review. *Hippocampus* **26**: 207–217. doi:10.1002/hipo.22561

Marr D. 1971. Simple memory: a theory for archicortex. *Philos Trans R Soc Lond B Biol Sci* **262**: 23–81. doi:10.1098/rstb.1971.0078

McEldoon JL, McNaughton BL, O’Reilly RC. 1995. Why there are complementary learning systems in the hippocampus and neocortex: insights from the successes and failures of connectionist models of learning and memory. *Psychol Rev* **102**: 419–457. doi:10.1037/0033-295X.102.3.419

McKenzie S, Frank AJ, Kinsky NR, Porter B, Riviere PD, Eichenbaum H. 2014. Hippocampal representation of related and opposing memories develop within distinct, hierarchically organized neural schemas. *Neuron* **83**: 202–215. doi:10.1016/j.neuron.2014.05.019

McNaughton BL, Morris RGM. 1987. Hippocampal synaptic enhacement and information storage within a distributed memory system. *Trends Neurosci* **10**: 408–415. doi:10.1016/0166-2236(87)90011-7

Milivojevic B, Varadinov M, Grabovsky AV, Collin SHP, Doeller CF. 2016. Coding of event nodes and narrative context in the hippocampus. *J Neurosci* **36**: 12412–12424. doi:10.1523/JNEUROSCI.2889-15.2016

Nakazawa K, Quirk MC, Chittwood WA, Watanabe M, Yeckel MF, Sun LD, Kato A, Carr CA, Johnston D, Wilson MA. 2002. Requirement for hippocampal CA3 NMDA receptors in associative memory recall. *Science* **297**: 211–218. doi:10.1126/science.1071795

Ngo CT, Horner AJ, Newcombe NS, Olson IR. 2019. Development of holistic episodic recollection. *Psychol Sci* **30**: 1696–1706. doi:10.1177/0956518119879441

Norman KA, O’Reilly RC. 2003. Modeling hippocampal and neocortical contributions to recognition memory: a complementary-learning-systems approach. *Psychol Rev* **110**: 641–666. doi:10.1037/1033-295X.110.4.611

O’Keefe J, Nadal L. 1978. The hippocampus as a cognitive map. Clarendon Press, Oxford.

Poole W, Gibbs DL, Shmulevich I, Bernard B, Knijnenburg TA. 2016. Combining dependent F-values with an empirical adaptation of Brown’s method. *Bioinformatics* **32**: i430–i436. doi:10.1093/bioinformatics/btw438

Rolls ET. 2015. Pattern separation, completion, and categorisation in the hippocampal and neocortical. *Neurol Learn Mem* **129**: 4–28. doi:10.1016/j.nlm.2015.07.008

Squire LR, Zola-Morgan S. 1991. The medial temporal lobe memory system. *Science* **238**: 1380–1386.

Stark SM, Stevenson R, Wu C, Rutledge S, Stark CEL. 2015. Stability of age-related deficits in the mnemonic similarity task across task variations. *Behav Neurosci* **129**: 257–268. doi:10.1037/bne0000055

Stokes J, Kyle C, Ekstrom A. 2013. Complementary roles of human hippocampal subfields in differentiation and integration of spatial context. *J Cogn Neurosci* **25**: 546–559. doi:10.1162/jocn_a_01857

Treves A, Rolls ET. 1992. Computational constraints suggest the need for two distinct input systems to the hippocampal CA3 network. *Hippocampus* **2**: 189–199. doi:10.1002/hipo.2209

Treves A, Tashiro A, Witter ME, Moser EI. 2016. What is the mammalian dentate gyrus for? *Front Aging Neurosci* **8**: 1–10. doi:10.3389/fnagi.2016.00008

Treves A, Tashiro A, Witter ME, Moser EI. 2008. What is the mammalian dentate gyrus for? *Front Aging Neurosci* **8**: 1–10. doi:10.3389/fnagi.2016.00008

Tulving E. 1983. *Elements of Episodic Memory*. Oxford University Press.

Uncapher MR, Rugg MD. 2005. Encoding and the durability of episodic memory: a functional magnetic resonance imaging study. *J Neurosci* **25**: 7260–7267. doi:10.1523/JNEUROSCI.1641-05.2005

Vazdarjanova A, Guzowski JF. 2004. Differences in hippocampal neuronal population responses to modifications in the context of the hippocampus: evidence for distinct, yet complementary, functions of CA3 and CA1 ensembles. *J Neurosci* **24**: 6489–6496. doi:10.1523/JNEUROSCI.0350-04.2004

Willa TJ, O’Reilly C, Cacucci F, Burgess N, Kneef J. 2005. Attractor dynamics in the hippocampal representation of the local environment. *Science* **308**: 873–876. doi:10.1126/science.1108905

Received April 13, 2020; accepted in revised form June 3, 2020.