Multi-scale Spectrum Sensing in Small-Cell mm-Wave Cognitive Wireless Networks

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Abstract—In this paper, a multi-scale approach to spectrum sensing in cognitive cellular networks is proposed. In order to overcome the huge cost incurred in the acquisition of full network state information, a hierarchical scheme is proposed, based on which local state estimates are aggregated up the hierarchy to obtain aggregate state information at multiple scales, which are then sent back to each cell for local decision making. Thus, each cell obtains fine-grained estimates of the channel occupancies of nearby cells, but coarse-grained estimates of those of distant cells. The performance of the aggregation scheme is studied in terms of the trade-off between the throughput achievable by secondary users and the interference generated by the activity of these secondary users to primary users. In order to account for the irregular structure of interference patterns arising from path loss, shadowing, and blockages, which are especially relevant in millimeter wave networks, a greedy algorithm is proposed to find a multi-scale aggregation tree to optimize the performance. It is shown numerically that this tailored hierarchy outperforms a regular tree construction by 60%.

I. INTRODUCTION

The recent proliferation of mobile devices has been exponential in number as well as heterogeneity [1]. This tremendous increase in demand as well as heterogeneity poses severe challenges due to the finite bandwidth of current systems, and calls for new tools for the design and optimization of agile wireless networks [2]. Cognitive radio [3] has the potential to improve spectral efficiency, by enabling smart terminals (secondary users, SUs) to exploit resource gaps left by legacy, primary users (PUs) [4].

In this paper, we consider a cognitive cellular network, which comprises a set of PUs, which are licensed to access the spectrum, and a set of SUs, which may access opportunistically any unoccupied spectrum. The network is arranged into cells. In each cell, PUs join and leave the channel at random times, thus the state of each cell is described by a first-order binary Markov process. In order to utilize the unoccupied spectrum, the SUs require accurate estimates of spectrum occupancies throughout the cellular network. In principle, the channel occupancies can be sensed locally in each cell and collected at a fusion center; the global network state information collected at the fusion center is then broadcasted to each cell for local decision making. In practice, however, such centralized estimation can be extremely costly in terms of transmit energy and time.

Due to path loss, shadowing, and blockage, SUs accessing the channel in one cell cause significant interference to nearby PUs, but negligible interference to distant PUs. Therefore, each SU needs precise information about the occupancies of nearby cells, but only coarse information about the occupancies of faraway cells. Given this intuition, we construct a cellular hierarchy, which is used to aggregate channel measurements over the network at multiple scales. Thus, SUs operating in a given cell have precise knowledge about the local state, aggregate knowledge of the states of the cells nearby, aggregate and coarser knowledge of the states of the cells farther away, and so on at multiple scales, reflecting the distance dependent nature of wireless interference.

This paper provides important extensions over [5], wherein we assumed a regular tree for hierarchical spectrum sensing by assuming that interference is regular and isotropic (matched to the hierarchy). Herein, we examine the irregular effects of shadowing and blockage, which are especially severe at millimeter wave frequencies [6]–[8]. As in [5], we tradeoff SU network throughput versus the interference generated by the SUs to the PUs. To overcome combinatorial complexity, we develop a greedy algorithm to determine the best hierarchical aggregation tree matched to the irregular interference patterns of millimeter wave communications. Optimality is defined in terms of the trade-off between SU network throughput and interference to PUs. As expected, this tailored hierarchy outperforms the regular tree construction in [5] by 60%. Our methods also apply to sub-6GHz wireless networks and are robust to issues of directionality of interferers and primary receivers.

Hierarchical estimation was proposed in [9] in the context of averaging consensus [10], which is a prototype for distributed, linear estimation schemes. Consensus-based schemes for spectrum estimation have also been proposed in [11], [12]. In contrast, we focus on a dynamic setting. A framework for joint spectrum sensing and scheduling in wireless networks has been proposed in [13], for the case of a single cell. Instead, we consider a network composed of multiple cells.

To summarize, the contributions of this paper are as follows. 1) We propose a hierarchical framework for aggregation of channel state information over a wireless network composed of multiple cells, with a generic interference pattern among cells. We study the performance of the aggregation scheme in terms of the trade-off between the throughput of SUs and the interference generated by the activity of the SUs to the PUs. 2) We develop a closed form expression for the belief of the spectrum occupancy vector that shows that this belief is statistically independent across subsets of cells at different levels of the hierarchy, and uniform within each subset (Theorem 1).
This results greatly facilitates the computations of the expected average long-term reward (Lemma 1): and 3) we address the optimal design of the hierarchical aggregation tree so as to optimize performance, for a given interference pattern; due to the combinatorial complexity of this problem, we propose a greedy algorithm based on agglomerative clustering [14, Ch. 14] (Algorithm 1).

This paper is organized as follows. In Sec. II, we present the system model. In Sec. III we present numerical results and, in Sec. VI, we conclude this paper. In Sec. IV, we present numerical results and, in Sec. VI, we conclude this paper.

II. SYSTEM MODEL

We consider a cognitive network, depicted in Fig. 1, composed of a primary cellular network with \( N_C \) cells, and an opportunistic network of SUs. Cells are indexed as \( 1, 2, \ldots, N_C \). We denote the set of cells as \( C \equiv \{ 1, 2, \ldots, N_C \} \). The SUs opportunistically access the spectrum so as to maximize their own throughput, under a constraint on the interference caused to the cellular network.

Let \( b_{i,t} \in \{ 0, 1 \} \) be the PU spectrum occupancy of cell \( i \in C \) in slot \( t \). That is, \( b_{i,t} = 1 \) if the channel is occupied by PUs in cell \( i \) at time \( t \), and \( b_{i,t} = 0 \) if it is idle. We suppose that \( \{ b_{i,t} \in \{ 0, 1 \} \}_{i 
 C} \) are i.i.d. across cells, and evolve according to a two-state Markov chain, as a result of PUs joining and leaving the network at random times. We let

\[
p = \mathbb{P}(b_{i,t+1} = 1 | b_{i,t} = 0), \quad q = \mathbb{P}(b_{i,t+1} = 0 | b_{i,t} = 1),
\]

be the transition probability of the Markov chain from "0" to "1" and from "1" to "0", respectively. Therefore, the steady-state probability that \( b_{i,t} \) is occupied is given by

\[
\pi_B \triangleq \frac{p}{p + q}.
\]

We denote the state of the network in slot \( t \) as \( b_t = (b_{1,t}, b_{2,t}, \ldots, b_{N_C,t}) \).

The activity of the SUs is represented by the SU access decision \( a_{i,t} \in \{ 0, 1 \} \), in cell \( i \), slot \( t \), where \( a_{i,t} = 1 \) if the SUs operating in cell \( i \) access the channel at time \( t \), and \( a_{i,t} = 0 \) otherwise. We denote the network-wide SU access decision as \( \mathbf{a}_t = (a_{1,t}, a_{2,t}, \ldots, a_{N_C,t}) \) in slot \( t \). The activity of the SUs generate interference to the cellular network. We denote the interference strength between cells \( i \) and \( j \) as \( \phi_{i,j} \geq 0 \). We assume that interference is symmetric, so that \( \phi_{i,j} = \phi_{j,i}, \forall i, j \in C \). Note that \( \phi_{i,j} \) is the strength of the interference caused by the SUs in cell \( i \) to cell \( j \). We let \( \Phi \) be the symmetric interference matrix, with components \( [\Phi]_{i,j} = \phi_{i,j}, \forall i, j \in C \).

Given the network state \( b_t \in \{ 0, 1 \}^{N_C} \) and the SU access decision \( a_{i,t} \in \{ 0, 1 \} \), we define the local reward for the SUs in cell \( i \) as

\[
r_i(a_{i,t}, b_t) = a_{i,t} r_i(1 - b_{i,t}) + \rho_B b_{i,t} - \lambda \sum_{j=1}^{N_C} \phi_{i,j} b_{j,t}.
\]

The term \( a_{i,t}(1 - b_{i,t}) \) in (3) equals one if and only if the SUs in cell \( i \) access the channel when cell \( i \) is idle; \( \rho_B \geq 0 \) is the instantaneous expected SU throughput accrued in this case. The term \( a_{i,t}b_{i,t} \) in (3) equals one if and only if the SUs in cell \( i \) access the channel when cell \( i \) is occupied; \( \rho_B \) is the instantaneous expected SU throughput accrued in this case, with \( 0 \leq \rho_B \leq \rho_I \). Finally, the term

\[
a_{i,t} \sum_{j=1}^{N_C} \phi_{i,j} b_{j,t}
\]

represents the overall interference generated by the SUs in cell \( i \) to the rest of the network, cell \( i \) included. The term \( \lambda > 0 \) is a Lagrangian multiplier which captures the trade-off between the reward for the SU system and the interference generated to the PUs.

The network reward is defined as the aggregate reward over the entire network, as a function of the SU access decision \( \mathbf{a}_t \) and network state \( b_t \),

\[
R(\mathbf{a}_t, b_t) = \sum_{i \in C} r_i(a_{i,t}, b_t).
\]

The SU access decision in cell \( i \) is decided based on partial network state information, denoted as \( \pi_{i,t} \) at time \( t \), where \( \pi_{i,t}(b) \) is the belief that the network state takes value \( b \). In slot \( t \), available to SUs in cell \( i \). Given \( \pi_{i,t} \), the SUs in cell \( i \) choose \( a_{i,t} \in \{ 0, 1 \} \) so as to maximize the expected reward

\[
r_i(a_{i,t}, \pi_{i,t}) \triangleq \sum_{b \in \{ 0, 1 \}^{N_C}} \pi_{i,t}(b) r_i(a_{i,t}, b).
\]

Thus,

\[
a_{i,t}^* \triangleq \arg \max_{a \in \{ 0, 1 \}} r_i(a, \pi_{i,t}),
\]

yielding the optimal expected local reward

\[
r_i^*(\pi_{i,t}) = \max\{ r_i(0, \pi_{i,t}), r_i(1, \pi_{i,t}) \} = \max\{ 0, r_i(1, \pi_{i,t}) \},
\]

where \( r_i(0, \pi_{i,t}) = 0 \) from (6).

Given the belief \( \pi_t = (\pi_{1,t}, \pi_{2,t}, \ldots, \pi_{N_C,t}) \) across the network, under the optimal SU access decisions \( \mathbf{a}_t \) given by (6), the optimal network reward is thus given by

\[
R^*(\pi_t) = \sum_{i \in C} r_i^*(\pi_{i,t}).
\]

Using the fact that \( r_i(a, b) \leq \max\{ r_i(0, b), r_i(1, b) \} \), \forall i \in C, \forall a \in \{ 0, 1 \}, \forall b \in \{ 0, 1 \}^{N_C} \), we obtain the inequality

\[
R^*(\pi_t) \leq \sum_{b \in \{ 0, 1 \}^{N_C}} \pi_{i,t}(b) \max\{ 0, r_i(1, b) \},
\]
As, the expected network reward under partial network state information is upper bounded by the expected network reward obtained when full network state information is provided to the SUs in each cell (perfect knowledge of $b_i$). Thus, the SUs should, possibly, obtain full network state information in order to achieve the best performance.

The belief $\pi_{i,t}$ is computed based on spectrum measurements performed over the network. Ideally, in order to achieve global network state information and maximize the reward (see [8]), the SUs in cell $i$ should obtain the local spectrum state $b_{i,t}$, as well as the spectrum state from the rest of the network. To this end, the SUs in cell $j \neq i$ should report the local spectrum state $b_{j,t}$ to the SUs in cell $i$ via information exchange, potentially over multiple hops, for transmitters/receivers far away from each other. Since this needs to be done over the entire network (i.e., for every pair $(i,j) \in \mathcal{C}^2$), the associated cost of information exchange may be huge, especially in large networks composed of a large number of small cells. In order to reduce the cost of acquisition of network state information, in this paper we propose a multi-scale approach to spectrum sensing. To this end, we structure the cellular grid in a hierarchical structure, defined by a tree of depth $D \geq 1$.

### A. Tree construction

Herein, we describe the tree construction. At level 0, we have the leaves, represented by the cells $\mathcal{C}$. We let $c_0^{(i)} \equiv \{i\}$ for $i \in \mathcal{C}$. At level 1, let $c_1^{(i)} = \{i\}$ for $i \in \mathcal{C}$, where $1 \leq n_1 \leq |\mathcal{C}|$. We associate a cluster head to each subset $\mathcal{C}_1^{(k)}$; the set of $n_1$ level-1 cluster-heads is denoted as $\mathcal{H}_1$. Hence, $\mathcal{C}_1^{(k)}$ is the set of cells associated to the level-1 cluster head $k \in \mathcal{H}_1$.

Recursively, at level $L$, let $\mathcal{H}_L$ be the set of nodes defining the level $L$-cluster heads, with $L \geq 1$. If $|\mathcal{H}_L| = 1$, then we have defined a tree with depth $D = L$. Otherwise, we define a partition of $\mathcal{H}_L$ into $n_{L+1}$ subsets $\mathcal{H}_L^{(i)}$, $k = 1, 2, \ldots, n_{L+1}$, where $1 \leq n_{L+1} \leq |\mathcal{H}_L|$, and we associate to each subset a level-$(L+1)$ cluster head; the set of $n_{L+1}$ level-$(L+1)$ cluster-heads is denoted as $\mathcal{H}_{L+1}$. Let $c_{L+1}^{(k)} = \{i\}$ for $i \in \mathcal{C}$ be the set of cells associated to level-$(L+1)$ cluster head $k \in \mathcal{H}_{L+1}$. This is obtained recursively as

$$c_{L+1}^{(k)} = \bigcup_{m \in \mathcal{H}_{L+1}^{(k)}} c_L^{(m)}. \quad (9)$$

Let $P_L(i) \in \mathcal{H}_L$ be the level $L$ parent of cell $i$, i.e., $P_0(i) = i$, and $P_L(i) = k$ for $L \geq 1$ if and only if $i \in c_k^{(L)}$, for some $k \in \mathcal{H}_L$. We make the following definition.

**Definition 1.** We define the hierarchical distance between cells $i \in \mathcal{C}$ and $j \in \mathcal{C}$ as

$$\Lambda(i,j) = \min \{L \geq 0 : P_L(i) = P_L(j)\}.$$  

In other words, $\Lambda(i,j)$ is the lowest level $L$ such that cells $i$ and $j$ belong to the same cluster at level $L$. It follows that $\Lambda(i,i) = 0$ and $\Lambda(i,j) = \Lambda(j,i)$, i.e., the hierarchical distance between cell $i$ and itself is 0, and it is symmetric.

We let $C^{(i)}(L)$ be the set of cells at hierarchical distance $L$ from cell $i$. That is, $C^{(i)}(0) = \{i\}$, and, for $L > 0$,

$$C^{(i)}(L) = C^{(P_L(i))} \setminus C^{(P_{L-1}(i))}. \quad (10)$$

In fact, by the tree construction, $C^{(P_L(i))}$ contains all cells with hierarchical distance (from cell $i$) less (or equal) than $L$. Thus, $C^{(i)}(L)$ is obtained by removing from $C^{(P_L(i))}$ all cells with hierarchical distance less (or equal) than $L − 1$, $C^{(P_{L-1}(i))}$.

### B. Hierarchical information exchange over the tree

In order to collect network state information at multiple scales, the SUs exchange local information over the tree. In particular, we propose a scheme in which the SUs carry out a hierarchical fusion of local estimates. This fusion is patterned after hierarchical averaging, a technique for scalar average consensus in wireless networks developed in [9].

At the beginning of slot $t$, at the cell level (level-0), the local SUs perform spectrum sensing to estimate the local state. Thus, the SUs in cell $i$ estimate the local state $b_{i,t}$ as $\hat{b}_{i,t}^{(i)} \in [0, 1]$, representing the belief that the local state takes the value $b_{i,t} = 1$, as seen by the SUs operating in cell $i$ (superscript $(i)$). For simplicity, in this paper we assume that local spectrum sensing is done with no errors, so that

$$\hat{b}_{i,t}^{(i)} = b_{i,t}, \forall i \in \mathcal{C}. \quad (11)$$

Next, these observations are fused up the hierarchy. The level 1 cluster head $m \in \mathcal{H}_1$ receives the spectrum measurements from its cluster $c_L^{(m)}$, and fuses them as

$$S_{m,t}^{(1)} = \sum_{j \in c_L^{(m)}} b_{j,t}, \forall m \in \mathcal{H}_1. \quad (12)$$

This process continues up the hierarchy: the level $L$ cluster head $m \in \mathcal{H}_L$ receives the aggregate spectrum measurements $S_{k,t}^{(L-1)}$ from the level-$(L-1)$ cluster heads $k \in \mathcal{H}_{L-1}$ connected to it, and fuses them as

$$S_{m,t}^{(L)} = \sum_{k \in \mathcal{H}^{(m)}_{L-1}} S_{k,t}^{(L-1)} = \sum_{j \in c_L^{(m)}} b_{j,t}, \forall m \in \mathcal{H}_L. \quad (13)$$

Eventually, the aggregate spectrum measurements are fused at the unique root of the tree (level $D$) as

$$S_{i,t}^{(D)} = \sum_{k \in \mathcal{H}^{(i)}_{D-1}} S_{k,t}^{(D-1)} = \sum_{j \in \mathcal{C}} b_{j,t}, \quad (14)$$

where we have used the fact that $C^{(D)} \equiv \mathcal{C}$.

Upon reaching level $D$, the appropriate aggregate spectrum measurements are propagated down to the individual cells $i \in \mathcal{C}$, following the tree. Thus, at the beginning of slot $t$, the SUs operating in cell $i$ receive

$$\left\{ \begin{array}{ll} S_{P_0(i),t}^{(0)} = \sum_{j \in c_0^{(P_0(i))}} b_{j,t} = b_{i,t}, \\
S_{P_L(i),t}^{(L)} = \sum_{j \in c_L^{(P_L(i))}} b_{j,t}, & 1 \leq L < D, \\
S_{i,t}^{(D)} = \sum_{j \in \mathcal{C}} b_{j,t}, & \end{array} \right. \quad (15)$$

where we remind that $P_L(i)$ is the level-$L$ parent of cell $i$, and $c_L^{(P_L(i))}$ is the set of cells associated to $P_L(i)$. That is, the SUs operating in cell $i$ receive from the level-$L$ parent the aggregate spectrum measurements over $C_L^{(P_L(i))}$. From this set of measurements, one can compute
Note that $\sigma_{i,t}^{(L)}$ is the aggregate spectrum measurement of the cells at hierarchical distance $L$ from cell $i$,
\begin{equation}
\sigma_{i,t}^{(L)} \triangleq \sum_{j \in C_{\Lambda}^{(L)}(L)} b_{j,t}, \forall 0 \leq L \leq D.
\end{equation}
Thus, the SUs in cell $i$ receive the set of aggregate spectrum measurements at multiple scales corresponding to different hierarchical distances. Importantly, they know only the aggregate spectrum measurements, but not the specific values of $b_{j,t}, \forall j \neq i$. These aggregate spectrum measurements are used to update the belief $\pi_{i,t}$ in the next section.

### III. Analysis

The SUs in cell $i$ update the belief $\pi_{i,t}$ based on past and present spectrum measurements $\sigma_{i,\tau} = (\sigma_{i,0}, \sigma_{i,1}, \ldots, \sigma_{i,D}), \forall 0 \leq \tau \leq t$. The form of $\pi_{i,t}$ is provided by the following Theorem.

**Theorem 1.** Given $\sigma_{i,t} = (o_0, o_1, \ldots, o_D)$,
\begin{equation}
\pi_{i,t}(b) = \prod_{l=0}^{D} \mathbb{P}(b_{j,t} = b_j, \forall j \in C_{\Lambda}^{(i)}(L) | \sigma_{i,t}^{(L)} = o_L),
\end{equation}

where the expectation is computed with respect to the state $s_{i,t}$.

**Proof.** See the Appendix.

From Equation (17), it follows that the belief $\pi_{i,t}$ is statistically independent across the subsets of cells at different hierarchical distances from cell $i$; this result follows from the fact that $\{b_{j,t}, t \geq 0, i \in C\}$ are i.i.d. across cells. Additionally, since $\sum_{j \in C_{\Lambda}^{(i)}(L)} b_{j,t} = o_L$ (as a result of state aggregation at hierarchical distance $L$) and $b_{j,t} \in \{0, 1\}$, there are\(\frac{|C_{\Lambda}^{(i)}(L)|}{D}\) possible combinations of $\{b_{j,t}, j \in C_{\Lambda}^{(i)}(L)\}$; Equation (18) states that these combinations are uniformly distributed, as a result of the i.i.d. assumption.

Importantly, $\pi_{i,t}$ is independent of past measurements but solely depends on the current one $\sigma_{i,t}$. In fact, spectrum occupancies $b_{j,t}$ are identically distributed across cells.

We can use Theorem 1 to compute the expected reward in cell $i$, given by (5). Using (5), we obtain
\begin{equation}
r_{i}(a_{i,t}, \pi_{i,t}) = \rho I_{a_{i,t}}(1 - \mathbb{P}(b_{i,t} = 1 | \pi_{i,t}))
+ \rho BA_{i,t} \mathbb{P}(b_{i,t} = 1 | \pi_{i,t}) - \lambda \sum_{j=1}^{N_C} \phi_{i,j} \mathbb{P}(b_{j,t} = 1 | \pi_{i,t})
= \rho I_{a_{i,t}}(1 - \mathbb{P}(b_{i,t} = 1 | \pi_{i,t})) + \rho BA_{i,t} \mathbb{P}(b_{i,t} = 1 | \pi_{i,t})
- \lambda \sum_{L=0}^{D} \sum_{j \in C_{\Lambda}^{(i)}(L)} \phi_{i,j} \mathbb{P}(b_{j,t} = 1 | \pi_{i,t}).
\end{equation}

In the last step above, we have partitioned the set of cells $C$ into the subsets corresponding to hierarchical distances $L = 0, 1, \ldots, D$ from cell $i$. Now, using (18) in Theorem 1 we obtain, for all $0 \leq L \leq D$, for all $\forall j \in C_{\Lambda}^{(i)}(L)$,
\begin{equation}
\mathbb{P}(b_{j,t} = 1 | \pi_{i,t}) = 0, \text{ if } o_L = 0,
\end{equation}
\begin{equation}
\mathbb{P}(b_{j,t} = 1 | \pi_{i,t}) = \frac{o_L \left( |C_{\Lambda}^{(i)}(L)| - o_L \right)}{|C_{\Lambda}^{(i)}(L)|} \frac{|C_{\Lambda}^{(i)}(L)| - 1}{o_L - 1}, \text{ if } o_L > 0,
\end{equation}
since there are $(|C_{\Lambda}^{(i)}(L)| - 1) o_L$ combinations such that $b_{j,t} = 1$, given that $\sigma_{i,t}^{(L)} = o_L$. Solving, we obtain
\begin{equation}
\mathbb{P}(b_{j,t} = 1 | \pi_{i,t}) = \frac{o_L}{|C_{\Lambda}^{(i)}(L)|}.
\end{equation}
Thus, substituting in (19), and letting
\begin{equation}
\Phi_{i}(L) \triangleq \sum_{j \in C_{\Lambda}^{(i)}(L)} \phi_{i,j}
\end{equation}
be the total interference generated by the SUs in cell $i$ to the cells at hierarchical distance $L$ from cell $i$, we can finally rewrite
\begin{equation}
r_{i}(a_{i,t}, \pi_{i,t}) = \rho I_{a_{i,t}}(1 - \pi_{i,t}^{(0)}) + \rho BA_{i,t} \pi_{i,t}^{(0)}
- \lambda \Phi_{i}(L),
\end{equation}
where, for convenience, we have expressed the dependence of $r_{i}(\cdot)$ on $\sigma_{i,t}$, rather than on $\pi_{i,t}$. Thus, the network reward $R_{i}$ is given by
\begin{equation}
R_{i}(\Sigma_{i}) = \max_{i \in C} \{0, r_{i}(1, \Sigma_{i})\},
\end{equation}
where we have defined $\Sigma_{i} = [\sigma_{i,t}, \sigma_{2,t}, \ldots, \sigma_{N_{C},t}]$, and, for convenience, we have expressed the dependence of $R_{i}(\cdot)$ on $\Sigma_{i}$, rather than on $\pi_{i}$.

#### A. Average long-term performance evaluation

We are interested in evaluating the average long-term performance of the hierarchical estimation scheme, that is
\begin{equation}
\bar{R} = \lim_{T \to \infty} \frac{1}{T} \mathbb{E} \left[ \sum_{t=0}^{T-1} R_{i}(\Sigma_{i}) \right],
\end{equation}
where the expectation is computed with respect to the sequence $\{\Sigma_{i}, t \geq 0\}$. We have the following result.

**Lemma 1.** The average long-term network reward is given by
\begin{equation}
\bar{R} = \sum_{i \in C} \sum_{a_{i,t} \in \{0, 1\}} \mathcal{B}(a_{i,t} | 1) \sum_{o_0 = 0}^{|C_{\Lambda}^{(i)}(1)|} \mathcal{B}(o_0 | C_{\Lambda}^{(i)}(1)) \cdots \sum_{o_{L-1} = 0}^{|C_{\Lambda}^{(i)}(L)|} \mathcal{B}(o_{L-1} | C_{\Lambda}^{(i)}(L)) \cdots \sum_{o_{D-1} = 0}^{|C_{\Lambda}^{(i)}(D)|} \mathcal{B}(o_{D-1} | C_{\Lambda}^{(i)}(D)) \times \max \left\{ \rho I_{a_{i,t}}(1 - o_0) + \rho B o_0 - \lambda \sum_{L=1}^{D} \frac{o_L}{|C_{\Lambda}^{(i)}(L)|} \Phi_{i}(L) \right\}.
\end{equation}
where \( B(h; |C_A^i(L)|) \) is the binomial distribution with \(|C_A^i(L)|\) trials and occupancy probability \( \pi_B \),
\[
B(h; |C_A^i(L)|) = \binom{|C_A^i(L)|}{h} \pi_B^h (1 - \pi_B)^{|C_A^i(L)| - h}.
\]

In fact, since channel occupancy states are i.i.d. across cells, at steady-state, the number of cells occupied within any subset \( \mathcal{C} \subseteq \mathcal{C} \) is a binomial random variable with \(|\mathcal{C}|\) trials (the number of cells in the set) and occupancy probability \( \pi_B \) (the steady-state probability that one cell is occupied). The result then follows by applying this argument to the hierarchical aggregation scheme.

Note that \( \bar{R} \) depends on the structure of the tree employed for network state information exchange. In the next section, we present an algorithm to design the tree so as to maximize the network reward \( \bar{R} \).

Using (9), we can compare the network reward \( \bar{R} \) with the upper bound computed under the assumption of full network state information at each cell, given by
\[
\bar{R}_{up} = \sum_{i \in \mathcal{C}} \sum_{b=0,1}^{NC} \pi_B^b (1 - \pi_B)^{NC - b} \left( \rho_i (1 - b_i) + \rho_B b_i - \lambda \sum_{j=1}^{NC} \phi_{i,j} b_j \right).
\]

IV. TREE DESIGN

The reward of the network depends crucially on the tree employed for information exchange. Optimizing the network reward over the set of all possible trees is a combinatorial problem with high complexity. Instead, we employ methods from hierarchical clustering to build a tree. Hierarchical clustering is well-studied (see, e.g. [14] Ch. 14), with two main approaches: divisive clustering, in which a tree is built by successively splitting larger clusters; and agglomerative clustering, in which a tree is built by successively combining smaller clusters. Our algorithm is based on the latter.

Agglomerative clustering requires a similarity metric between clusters; at each round, similar clusters are aggregated. Our goal in designing a tree-based approach to spectrum sensing is to prioritize information that nodes can use to limit the interference they generate to other cells. Therefore, we want to aggregate cells together with high potential for interference. To this end, we define the similarity between level-\(L\) clusters \( k_1, k_2 \in \mathcal{H}_L \) as
\[
\gamma_L(k_1, k_2) = \sum_{i \in C_{k_1}^L} \sum_{j \in C_{k_2}^L} \phi_{i,j},
\]
the sum of inter-cluster interference strengths.

The algorithm proceeds as shown in Algorithm 1. We initialize it with the \( NC \) leaves \( C_0^i = \{i\}, i = 1, 2, \ldots, NC \). Then, at each level \( L \), we iterate over all of the clusters, pairing each one with the cluster with which it has the most interference (this can be done in order with pairs with maximum similarity [29]). This forms the set of level \( L+1 \) clusters. If the number of clusters at level \( L \) happens to be odd, one cluster may not be paired, in which case it forms its own cluster at level \( L + 1 \). The algorithm continues until the cluster \( C_{L+1} \) contains the entire network, i.e., a tree is formed.

Agglomerative clustering has complexity \( O(N_C^2 \log(N_C)) \), where the \( N_C^2 \) term owes to searching over all pairs of clusters.

V. NUMERICAL RESULTS

In this section, we provide numerical results. We consider a \( 4 \times 4 \) cells network. We set the parameters as follows: \( \rho_I = 1, \quad \rho_B = 0, \quad \lambda = 1, \quad p = q = 0.1 \). We use the following interference model between a pair of cells (assuming there is no blockage between them):
\[
\left\{ \begin{array}{ll}
\phi_{i,j} = \|p(i) - p(j)\|^{-\alpha}, & i \neq j, \\
\phi_{i,i} = 1,
\end{array} \right.
\]

\[\text{(30)}\]
where \( p(i) \) is the position of cell \( i \), \( \| p(i) - p(j) \| \) is the distance between cells \( i \) and \( j \), and \( \alpha = 2 \) is the pathloss exponent.

We define random "walls" between cell boundaries, i.i.d. with probability \( p_{\text{block}} \in [0, 1] \). If a wall is present, then all the cells separated by it experience mutual blockage; thus, if cells \( i \) and \( j \) are separated by a wall, then \( \phi_{i,j} = \phi_{j,i} = 0 \). The blockage probability is generated randomly, for a given blockage probability \( p_{\text{block}} \), and a sample average of the performance is computed over 200 independent trials.

In Fig. 2 we plot the curve of the network reward as a function of the blockage probability \( p_{\text{block}} \), for different schemes:

- a scheme in which a regular tree is used for state information aggregation. In this case, neighboring cells and clusters are paired together, in order, independently of the interference pattern. This scheme is similar to [5];
- a scheme in which the tree is generated with Algorithm 1 by leveraging the specific structure of interference;
- an upper bound in which the reward is computed under full network state information at each cell, given by (28).

This is computed via Monte Carlo simulation over 5000 independent realizations of \( b_t \) (at steady-state).

We notice that the best performance is obtained under full network state information available at each cell. This is because each cell can leverage the most refined information on the interference pattern. However, this comes at a huge cost of propagating network state information over the network. In contrast, the cost of acquisition of state information can be significantly reduced using aggregation, at the cost of some performance degradation. Remarkably, by using the greedily optimized algorithm for information aggregation, the performance improves by up to 60% with respect to a scheme that uses a regular tree. In fact, the greedily optimized algorithm leverages the specific structure of interference over the network.

![Fig. 2: Average long-term network reward as a function of the probability of blockage \( p_{\text{block}} \).](image)

VI. CONCLUSIONS

In this paper, we have proposed a multi-scale approach to spectrum sensing in cognitive cellular networks. To reduce the cost of acquisition of network state information, we have proposed a hierarchical scheme, that makes it possible to obtain aggregate state information at multiple scales, at each cell. We have studied analytically the performance of the aggregation scheme in terms of the trade-off between the throughput achievable by secondary users and the interference generated by the activity of these secondary users to primary users. We have proposed a greedy algorithm to find a multi-scale aggregation tree, matched to the structure of interference, to optimize the performance. Finally, we have shown performance improvement up to 60% using a greedily optimized tree, compared to a regular one.

APPENDIX: PROOF OF THEOREM 1

**Proof.** We prove (17) and (18) by induction on \( t \). At time \( t = 0 \), given \( \sigma_{i,0} = (o_0, o_1, \ldots, o_D) \), using Bayes’ rule we obtain

\[
\pi_{i,0}(b) = \mathbb{P}(b_{j,0} = b_j, \forall j | \sigma_{i,0} = (o_0, o_1, \ldots, o_D)) \tag{31}
\]

\[
= \frac{\sum_{b_{j,0} \in \{0,1\}^NC} \mathbb{P}(\sigma_{i,0} = (o_0, o_1, \ldots, o_D) | b_{j,0} = b_j, \forall j) \mathbb{P}(b_{j,0} = b_j, \forall j)}{\mathbb{P}(\sigma_{i,0} = (o_0, o_1, \ldots, o_D))}
\]

Then, noticing that \( \sigma_{i,0} \) in (16) is only a function of \( b_{j,0}, \forall j \in C_{A}^{(i)}(L) \), but is independent of \( b_{j,0}, j \notin C_{A}^{(i)}(L) \), and since \( b_{j,0} \) is statistically independent across cells \( j \), we obtain

\[
\pi_{i,0}(b) = \left[ \prod_{L=0}^{D} \mathbb{P}(\sigma_{i,L}^{(L)} = o_L | b_{j,0} = b_j, \forall j \in C_{A}^{(i)}(L)) \times \mathbb{P}(b_{j,0} = b_j, \forall j) \right].
\]

Using the fact that \( \{C_{A}^{(i)}(L), L = 0, 1, \ldots, D\} \) define a partition of \( \{1, 2, \ldots, N_C\} \), we have that

\[
\sum_{b} \prod_{L=0}^{D} f_L(\tilde{b}_j, j \in C_{A}^{(i)}(L)) = \prod_{L=0}^{D} \sum_{b_{j,0} \in C_{A}^{(i)}(L)} f_L(\tilde{b}_j, j \in C_{A}^{(i)}(L)),
\]

for generic functions \( f_L : \{0, 1\}^{C_{A}^{(i)}(L)} \mapsto \mathbb{R} \), hence by using this fact in the denominator of (32) we obtain

\[
\pi_{i,0}(b) = \left[ \prod_{L=0}^{D} \mathbb{P}(\sigma_{i,L}^{(L)} = o_L | b_{j,0} = \tilde{b}_j, \forall j \in C_{A}^{(i)}(L)) \times \mathbb{P}(b_{j,0} = \tilde{b}_j, \forall j) \right].
\]
By Bayes’ rule we finally obtain
\[
\pi_{i,0}(b) = \prod_{L=0}^{D} \mathbb{P}(b_{j,0} = b_j, \forall j \in C_{\Lambda}^{(i)}(L)|\sigma_{i,0}^{(L)} = o_L),
\]
yielding (17) for \( t = 0 \).

Note that
\[
\mathbb{P}(b_{j,0} = b_j, \forall j \in C_{\Lambda}^{(i)}(L)|\sigma_{i,0}^{(L)} = o_L) = \frac{\sum_{b_j \in \{0,1\}} \mathbb{P}(\sigma_{i,0}^{(L)} = o_L|b_{j,0} = b_j, \forall j \in C_{\Lambda}^{(i)}(L)) \times \mathbb{P}(b_{j,0} = b_j, \forall j \in C_{\Lambda}^{(i)}(L))}{\sum_{b_j \in \{0,1\}} \mathbb{P}(\sigma_{i,0}^{(L)} = o_L|b_{j,0} = \tilde{b}_j, \forall j \in C_{\Lambda}^{(i)}(L)) \times \mathbb{P}(b_{j,0} = \tilde{b}_j, \forall j \in C_{\Lambda}^{(i)}(L))},
\]
then, using the definition of \( \sigma_{i,0}^{(L)} \) in (16), and noticing that it is a function of \( b_{j,0}, \forall j \in C_{\Lambda}^{(i)}(L) \), we obtain
\[
\mathbb{P}(b_{j,0} = b_j, \forall j \in C_{\Lambda}^{(i)}(L)|\sigma_{i,0}^{(L)} = o_L) = \frac{\chi \left( \sum_{j \in C_{\Lambda}^{(i)}(L)} b_j = o_L \right)}{\prod_{j \in C_{\Lambda}^{(i)}(L)} \mathbb{P}(b_{j,0} = b_j)} \prod_{j \in C_{\Lambda}^{(i)}(L)} \mathbb{P}(b_{j,0} = \tilde{b}_j),
\]
where \( \mathbf{b} = (b_{j,0}) \)

Note that, if \( \mathbf{b} \) is such that \( \sum_{j \in C_{\Lambda}^{(i)}(L)} b_j \neq o_L \), then
\[
\mathbb{P}(b_{j,0} = b_j, \forall j \in C_{\Lambda}^{(i)}(L)|\sigma_{i,0}^{(L)} = o_L) = 0
\]
as in (18). Conversely, if \( \sum_{j \in C_{\Lambda}^{(i)}(L)} b_j = o_L \), for all vectors \( \mathbf{b} \) such that \( \sum_{j \in C_{\Lambda}^{(i)}(L)} \tilde{b}_j = o_L \) we have that
\[
\prod_{j \in C_{\Lambda}^{(i)}(L)} \mathbb{P}(b_{j,0} = b_j) = \prod_{j \in C_{\Lambda}^{(i)}(L)} \mathbb{P}(b_{j,0} = \tilde{b}_j),
\]
since \( b_{j,0} \) are identically distributed. Thus it follows that
\[
\mathbb{P}(b_{j,0} = b_j, \forall j \in C_{\Lambda}^{(i)}(L)|\sigma_{i,0}^{(L)} = o_L) = \frac{\chi \left( \sum_{j \in C_{\Lambda}^{(i)}(L)} b_j = o_L \right)}{\prod_{j \in C_{\Lambda}^{(i)}(L)} \mathbb{P}(b_{j,0} = b_j)} \prod_{j \in C_{\Lambda}^{(i)}(L)} \mathbb{P}(b_{j,0} = \tilde{b}_j),
\]
when
\[
\mathbf{b} = (b_{j,0}) \]

Note that, since \( \{b_{j,t}, t \geq 0\} \) is a Markov chain, we obtain
\[
\mathbb{P}(b_{j,t} = b_j, \forall j \mid \sigma_{i,t}, \tau \leq t - 1) = \sum_{\mathbf{b}} \mathbb{P}(b_{j,t} = b_j, \forall j \mid \sigma_{i,t}, \tau \leq t - 1)
\]
\[
= \prod_{L=0}^{D} \sum_{b_{j,t-1} = \tilde{b}_j} \mathbb{P}(b_{j,t} = b_j, \forall j \mid \sigma_{i,t-1} = \tilde{b}_j),
\]
where we used (33). Then, using (18) to express \( \mathbb{P}(b_{j,t-1} = \tilde{b}_j, \forall j \in C_{\Lambda}^{(i)}(L)|\sigma_{i,t-1} = o_L) \) and substituting the resulting
\[
\mathbb{P}(b_{j,t} = b_j, \forall j \mid \sigma_{i,t}, \tau \leq t - 1) = \sum_{\mathbf{b}} \mathbb{P}(b_{j,t} = b_j, \forall j \mid \sigma_{i,t}, \tau \leq t - 1)
\]
show that they hold at time \( t \) as well. We have
\[
\pi_{i,t}(b) = \mathbb{P}(b_{j,t} = b_j, \forall j \mid \sigma_{i,t}, \tau \leq t)
\]
\[
= \prod_{L=0}^{D} \sum_{b_{j,t-1} = \tilde{b}_j} \mathbb{P}(b_{j,t} = b_j, \forall j \mid \sigma_{i,t}, \tau \leq t - 1)
\]
\[
= \prod_{L=0}^{D} \sum_{b_{j,t-1} = \tilde{b}_j} \mathbb{P}(b_{j,t} = b_j, \forall j \mid \sigma_{i,t-1} = \tilde{b}_j),
\]
where we used (33). Then, using (18) to express \( \mathbb{P}(b_{j,t-1} = \tilde{b}_j, \forall j \in C_{\Lambda}^{(i)}(L)|\sigma_{i,t-1} = o_L) \) and substituting the resulting
expression in (42), we obtain
\[\pi_{i,t}(b) = \prod_{L=0}^{D} \frac{\sum_{b,\hat{b}} \chi \left( \sum_{j \in C_{A}^{(i)}(L)} b_j = o_L, \sum_{j \in C_{A}^{(i)}(L)} \hat{b}_j = \hat{o}_L \right) \prod \mathbb{P}(b_j, t = b_j | b_{j,t-1} = \hat{b}_j)}{\prod_{j \in C_{A}^{(i)}(L)} \mathbb{P}(b_j, t = b_j | b_{j,t-1} = \hat{b}_j)}\]  
(50)

Note that, for any \(b\) and \(\hat{b}\), and for any permutation \(P(\cdot) : C_{A}^{(i)}(L) \to C_{A}^{(i)}(L)\) of the elements in the set \(C_{A}^{(i)}(L)\), we have
\[\prod_{j \in C_{A}^{(i)}(L)} \mathbb{P}(b_j, t = b_j | b_{j,t-1} = \hat{b}_j) = \prod_{j \in C_{A}^{(i)}(L)} \mathbb{P}(b_{P(j)}, t = b_{P(j)} | b_{P(j),t-1} = \hat{b}_{P(j)}),\]  
(51)

since \(b_{j,t}\) are statistically identical across cells.

Thus, for any \(b\) and \(\tilde{b}\) such that
\[\sum_{j \in C_{A}^{(i)}(L)} b_j = \sum_{j \in \tilde{C}_{A}^{(i)}(L)} \tilde{b}_j = o_L,\]  
(53)

we obtain
\[\prod_{j \in C_{A}^{(i)}(L)} \mathbb{P}(b_j, t = b_j | b_{j,t-1} = \hat{b}_j) = \prod_{j \in \tilde{C}_{A}^{(i)}(L)} \mathbb{P}(b_j, t = \tilde{b}_j | b_{j,t-1} = \tilde{b}_{P(j)}),\]  
(54)

Finally, substituting in (50) we obtain
\[\pi_{i,t}(b) = \prod_{L=0}^{D} \frac{\sum_{b,\hat{b}} \chi \left( \sum_{j \in C_{A}^{(i)}(L)} b_j = o_L, \sum_{j \in C_{A}^{(i)}(L)} \hat{b}_j = \hat{o}_L \right) \prod \mathbb{P}(b_j, t = b_j | b_{j,t-1} = \hat{b}_j)}{\prod_{j \in C_{A}^{(i)}(L)} \mathbb{P}(b_j, t = b_j | b_{j,t-1} = \hat{b}_j)}\].

Note that this expression implies that \(\pi_{i,t}(b)\) is only a function of \(\sigma_{i,t}\), but is independent of \(\sigma_{i,\tau} \tau \leq t-1\); additionally, \(\{b_{j}, j \in C_{A}^{(i)}(L)\}\) is statistically independent of \(\{b_{j}, j \in C_{A}^{(i)}(L)\}\) for \(L \neq L_2\), given \(\sigma_{i,t}\). Thus, (17) follows, where \(\mathbb{P}(b_{j,t} = b_j, \forall j \in C_{A}^{(i)}(L) | \sigma_{i,t}^{(L)} = o_L)\) is given as in (39).

Finally, (18) follows from (39), (40).

The induction step, and the Theorem, are thus proved. □

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