INTRODUCTION

Gliomas are the most common occurring type of brain tumors worldwide, accounting for approximately 80% of all malignant brain tumors. In 2000, it is estimated that the annual incidence of brain tumor was about 3.9/105 in men and 2.8 per 105 in women in Chinese population, and it is about 2–3 per 100,000 in Europe and North America. The incidence of brain tumors has increased in recent several years, and the incidence shows significant different in terms of gender, age, race, ethnicity, and even geographical region. However, the real etiology of glioma is still not well understood. Previous studies reported that mutations of many genes can affect the risk of glioma.
It is well known that changes in DNA induced by exposure to environmental and endogenous agents may result in genetic instability and mutagenesis. DNA repair mechanisms play an important role in maintaining genomic integrity and preventing carcinogenesis. Therefore, if genetic variations in DNA repair genes change the amount or activity of the proteins that they encode, they might also increase the risk of disease and the risk of cancer in particular. Excision repair cross complementation group 1 (ERCC1) and ERCC2 lie on chromosome 19q13.3, and they play an important role in NER pathway.10

Several previous studies have reported the role variant of ERCC1 rs3212986 and ERCC2 rs13181 polymorphisms in the risk of glioma, but the results of these studies are inconsistent. Therefore, we aimed to conduct a meta-analysis to investigate the role of ERCC1 rs3212986 and ERCC2 rs13181 on the risk of glioma.

METHODS

Search strategy: A comprehensive research was conducted through the databases of Pubmed, EMBASE and the China National Knowledge Infrastructure (CNKI) platforms until June 1, 2014. The literature search was conducted using the following terms: “glioma” or “brain tumor”, “polymorphism” or “variant”, and “ERCC1 C8092A”, “ERCC2 Lys751Gln”, “rs3212986” or “rs13181”. The reference lists of articles included for review and past meta-analyses were examined for any further relevant publications. No publication date or language restrictions were applied.

Study selection: The inclusion criteria for studies were as follows: studies are case-control design; studies evaluated the association between ERCC1 rs3212986 and ERCC2 rs13181 polymorphisms and glioma risk; studies reported the results of available genotype frequencies.

The exclusion criteria for studies were as follows: articles only have an abstract, review articles and comments; studies are overlapped with other studies; studies have no comparison or control group; studies have no data of genotype frequencies. Corresponding authors were contacted in an attempt to obtain unreported genotype counts if studies were otherwise eligible.

Data extraction
Two authors independently screened the electric search with all terms. Duplications and obviously irrelevant studies were excluded according to the exclusion criteria. We extracted the full texts of all eligible studies according to the inclusion criteria. The study ID, study design, participation number, case-control number and genotype information were extracted for each eligible study.

Statistical analysis: All meta-analysis analysis was conducted by STATA 9.0 software. A chi-square ($\chi^2$) was taken to evaluate the Hardy–Weinberg equilibriums in control groups. The pooled odds ratios (OR) and 95% confidence interval (CI) were taken to calculate the role of ERCC1 rs3212986 and ERCC2 rs13181 polymorphisms on the risk of glioma. The heterogeneity between studies was estimated by $I^2$ test and heterogeneity Q statistic test. When $I^2$ were at the range of 0-25%, there was no degree of heterogeneity. When $I^2$ were at the range of 25-50%, there was moderate heterogeneity. When $I^2$ were at the range of 50-75%, there was high heterogeneity. When $I^2$ were at the range of 75-100%, there was extreme heterogeneity.

Table I: Meta-analysis of the association of ERCC1 rs3212986 polymorphism with the glioma risk

| Study ID | Country  | Cases | Total cases | Controls | Total controls | $P$ for HWE | OR(95%CI)$^a$ |
|----------|----------|-------|-------------|----------|----------------|-------------|--------------|
|          | CC       | CA    | AA          | CC       | CA             |             |              |
| Chen P   | 2000     | Caucasian | 73   | 43       | 122  | 81  | 70  | 159 | 0.47 | 0.68(0.42-1.12) 0.83(0.28-2.51) |
| Wrensch M| 2005     | Caucasian | 206  | 144      | 259  | 237 | 184 | 410 | 0.09 | 0.90(0.68-1.20) 1.25(0.69-2.27) |
| Liu Y    | 2009     | Caucasian | 208  | 130      | 33   | 219 | 126 | 302 | 0.83 | 1.09(0.80-1.48) 1.92(1.03-3.57) |
| McKean- Cowdin R    | 2009 | Caucasian | 557  | 361      | 59   | 1087 | 728 | 105 | 362 | 0.24 | 0.97(0.82-1.14) 1.10(0.78-1.53) |
| Chen DQ  | 2012     | Chinese  | 202  | 141      | 50   | 977  | 221 | 154 | 35  | 1.00 | 1.00(0.74-1.35) 1.56(0.97-2.51) |
| Zhang N  | 2012     | Chinese  | 123  | 98       | 36   | 443  | 144 | 105 | 29  | 0.14 | 1.09(0.76-1.38) 1.45(0.84-2.51) |
| Pan WR   | 2013     | Chinese  | 229  | 169      | 45   | 375  | 241 | 162 | 41  | 0.08 | 0.10(0.83-1.41) 1.16(0.73-1.83) |
| Dong YS  | 2014     | Chinese  | 33   | 32       | 7    | 257  | 137 | 144 | 21  | 0.94 | 0.92(0.54-1.58) 1.38(0.54-3.53) |
| Pooled results | 1631 | 1118 | 259 | 3008 | 2367 | 1673 | 279 | 4319 | 0.98 | 0.98(0.89-1.09) 1.29(1.07-1.55) |

$^a$OR: odds ratio; CI: confidence interval; HWE: Hardy–Weinberg equilibrium. $P < 0.05$, it was considered statistically significant.
the range of 75-100%, there was great heterogeneity. A random-effects model or fixed-effect model was taken to calculate the pooled OR (95% CI) according to the degree of heterogeneity between studies. The publication bias in studies was calculated using Begg’s funnel plot and Egger’s test.

RESULTS

Characteristics of eligible publications: Our comprehensive literature search identified a total of 57 studies for ERCC1 rs3212986 and ERCC2 rs13181 based on their titles. Finally, 14 eligible case-control studies were selected, including eight studies for ERCC1 C8092A and 11 studies for ERCC2 Lys751Gln.11-24

Eight studies reported the association between ERCC1 rs3212986 polymorphism and glioma risk, including 3008 glioma cases and 4319 controls (Table-I).11-18 11 studies reported the association between ERCC2 Lys751Gln polymorphism and glioma risk, including 3456 glioma cases and 4957 controls (Table-II).12-15,19-25

Six studies were conducted in Chinese population, and the other seven studies were conducted in Caucasian populations.

The genetic distributions of ERCC1 rs3212986 in eight studies and ERCC2 rs13181 in nine studies were in according to Hardy-Weinberg Equilibrium (HWE). Only one study about ERCC2 rs13181 deviated from Hardy-Weinberg equilibrium in the control groups.

There was no significant heterogeneity between pooled studies, and thus we performed a fixed-effect model to assess the association between ERCC1 rs3212986 and ERCC2 rs13181 and risk of glioma. Our meta-analysis found that ERCC1 8092 AA genotype was significantly associated with increased risk of glioma compared with CC genotype, and the pooled OR (95% CI) was 1.29 (1.07-1.55). However, we did not find significant association between ERCC2 rs13181 polymorphisms and risk of glioma. By subgroup analysis, ERCC1 rs3212986 AA genotype was found to be significantly correlated with increased glioma risk in Chinese population (OR=1.37, 95% CI=1.07, 1.55) (Fig.1). Similarly, we found that ERCC2 rs13181 GT and TT genotypes were significantly associated with increased risk of glioma.

OR: odds ratio; CI: confidence interval; HWE: Hardy–Weinberg equilibrium. aP < 0.05, it was considered statistically significant.

![Fig.1: Subgroup analysis for the role of ERCC1 rs3212986 AA on the glioma risk.](image-url)
of glioma in Chinese population, with ORs (95% CI) of 1.47 (1.17-1.85) and 1.50 (1.02-2.22). But ERCC1 rs3212986 and ERCC2 rs13181 polymorphisms had no significant association with glioma risk in Caucasian populations (Fig. 2 and Fig. 3).

By begg’s funnel plot, we found that the shapes of the funnel plots for ERCC1 rs3212986 and ERCC2 rs13181 were symmetry, which suggests no publication bias in this meta-analysis (Fig. 4 and Fig. 5). Moreover, sensitivity analysis showed that the pooled OR did not alter after excluding one largest sample size study.

**DISCUSSION**

To the best of our knowledge, the current meta-analysis comprehensively analyze the association between ERCC1 rs3212986 and ERCC2 rs13181 polymorphism and glioma risk. The results of this meta-analysis suggest that the ERCC1 rs3212986 and ERCC2 rs13181 polymorphism polymorphisms was significantly associated with increased risk of glioma in Chinese population, but no association in Caucasian Chinese.

ERCC1 and ERCC2 genes are two important rate-limiting enzymes, and they play a role in the NER process. ERCC1 is a subunit of the NER complex, and ERCC1 usually interacts with XPA, XPF and RPA and can guides 5’ cleavage activity during the process of NER. Deficient of ERCC1 expression usually shows a high mutation frequency, an increased level of genomic instability and decreased S-phase-dependent illegitimate chromosome exchange, and it can cause a response adopted by rodent cells to prevent the accumulation of DNA double strand breaks. ERCC2 is located at chromosome 19q13.3, and this gene possesses both single strand DNA-dependent ATPase and 5’-3’ DNA helicase activities and participates in DNA unwinding during the process of NER pathway.
Variants of ERCC2 gene is reported to have a role in decreasing the helicase activity, causing a lower DNA repair capacity of NER pathway and affecting the risk of cancer.\(^{31,32}\) Therefore, variants of ERCC1 and ERCC2 genes can influence the cellular DNA repair capacity and affect the susceptibility to glioma risk.\(^{11,13}\) Our meta-analysis demonstrated that ERCC1 rs3212986 polymorphism was significantly associated with glioma risk, but ERCC2 rs13181 polymorphism was not correlated with glioma risk. When conducting stratified analysis by ethnicity, ERCC1 rs3212986 AA, ERCC2 rs13181 GT and TT genotypes was associated with increased risk of glioma in Chinese population. But there was no significant increased risk in Caucasian population. Our study suggests that ERCC1 rs3212986 and ERCC2 rs13181 showed different effect on the risk of glioma across different populations.

In our study, we did not detect significant difference between-study heterogeneity of all included case-control studies, and the pooled OR did not alter using the sensitivity analysis. Our study indicates that the results of our study are robust and high stable.

Several limitations in our study should be considered. First, the sample size of cases and controls in the stratified analysis was relatively small, which could reduce the statistical power to find the difference between groups. Second, the studies involving Asian patients were too limited to identify precise gene associations. Third, we could not perform a stratified analysis according to ethnicity because of insufficient data. Future studies may address the effect of the ERCC1 rs3212986 and ERCC2 rs13181 polymorphism on different histological brain tumor subtypes to provide a better understanding of tumor pathogenesis.

In summary, our meta-analysis suggested that ERCC1 rs3212986 and ERCC2 rs13181 polymorphism play an important risk factor for brain tumor development in Chinese population, but has no association in Caucasian populations. Since potential biases and confounders could not be completely excluded from this meta-analysis, further studies with larger sample sizes and standardized, unbiased genotyping methods are warranted to confirm our findings.

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