THE X-RAY POWER SPECTRAL DENSITY FUNCTION OF THE SEYFERT ACTIVE GALACTIC NUCLEUS NGC 7469

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ABSTRACT

We present the broadband X-ray power spectral density (PSD) function of the X-ray-luminous Seyfert 1.2 NGC 7469, measured from Rossi X-ray Timing Explorer monitoring data and two XMM-Newton observations. We find significant evidence for a turnover in the 2–10 keV PSD at a temporal frequency of $2.0^{+3.0}_{−0.8} \times 10^{-6}$ Hz or $1.0^{+3.0}_{−0.6} \times 10^{-6}$ Hz, depending on the exact form of the break (sharp break or slowly bending power law, respectively). The “surrogate” Monte Carlo method of Press et al. was used to map out the probability distributions of PSD model parameters and obtain reliable uncertainties (68% confidence limits quoted here). The corresponding break timescales of $5.8 \pm 3.5$ days or $11.6^{+17.5}_{−5}$ days, respectively, is consistent with the empirical relation between PSD break timescale, black hole mass, and bolometric luminosity of McHardy et al. Compared to the 2–10 keV PSD, the 10–20 keV PSD has a much flatter shape at high temporal frequencies, and no PSD break is significantly detected, suggesting an energy-dependent evolution not unlike that exhibited by several Galactic black hole systems.

Key words: galaxies: active – galaxies: Seyfert – X-rays: individual (NGC 7469)

Online-only material: color figures

1. INTRODUCTION

In Seyfert active galactic nuclei (AGNs), the observation of ubiquitous rapid variability of the X-ray continuum emission has established that these X-rays must originate very close to the central supermassive black hole, in the central tens of Schwarzschild radii ($R_{\text{Sch}} \equiv 2GM_\text{BH}/c^2$), and in the region of extreme gravity. The X-rays are generally thought to originate in some type of optically thin, hot ($10^9$ K) corona; plausible geometrical configurations include, but are not limited to, the hot, ionized outer layers of the optically thick, geometrically thin accretion disk which feeds the black hole (Nayakshin et al. 2000), part of a central, advection-dominated, radiatively inefficient accretion flow (e.g., Narayan & Yi 1994) or the base of an outflowing jet (Markoff et al. 2005).

A common tool to quantify the aperiodic variability as a function of temporal frequency is the power spectral density (PSD) function. The Rossi X-ray Timing Explorer (RXTE), launched in 1995, has allowed multi-timescale monitoring campaigns to be executed, yielding high-quality, broadband X-ray PSDs probing X-ray variability on temporal frequencies from $10^{-4}$ to $10^{-8}$ Hz (timescales from hours to years). These PSDs commonly show “breaks” at temporal frequencies $f_b$, with PSD power-law slopes breaking from $-2$ to $-1$ above and below $f_b$, respectively. Values of $f_b$ are usually in the range $10^{-6}$ to $10^{-3}$ Hz, i.e., break timescales $T_b \equiv 1/f_b$ are usually in the range $\sim 0.01$ to a few days (e.g., Edelson & Nandra 1999; Uttley et al. 2002, hereafter U02; Markowitz et al. 2003, hereafter M03; McHardy et al. 2004, 2005, 2007; Markowitz et al. 2007; Porquet et al. 2007; Chatterjee et al. 2009). These papers also demonstrated that break frequencies were consistent with a correlation between $T_b$ and black hole masses $M_\text{BH}$, estimated by optical measurements such as reverberation mapping (e.g., Peterson et al. 2004) or via host galaxy stellar velocity dispersion $\sigma_v$ (e.g., Merritt & Ferrarese 2001; Tremaine et al. 2002). The correlation extrapolates down to stellar-mass, Galactic black hole X-ray binary (BH XRB) systems, and, combined with the remarkable similarity in PSD shapes between Seyferts and BH XRBs, strongly supports the notion of identical variability mechanisms operating in both classes of objects.

However, as first suggested by McHardy et al. (2004), there is an additional dependence on the accretion relative to Eddington $m \equiv L_{\text{bol}}/L_{\text{Edd}}$, such that for a given $M_\text{BH}$, $T_b$ decreases as $m$ increases. This “fundamental plane” between $T_b$, $M_\text{BH}$, and $m$ signifies a sort of “unification” across supermassive and stellar-mass accreting black hole systems, and has been quantified in an empirically derived relation by McHardy et al. (2006) from a sample of Seyferts and BH XRB X-ray PSDs.

In this paper, we present the broadband X-ray PSD of the Seyfert AGN NGC 7469 and present evidence for a PSD break at timescale of $\lesssim 10$ days. As this source’s black hole mass is well known both from reverberation mapping the optical broad-line region (from an International AGN Watch monitoring program, e.g., Peterson et al. 2004) and from the $M_\text{BH}–\sigma_v$ relation, we can test if the break is consistent with the McHardy et al. (2006) empirical relation. Nandra & Papadakis (2001, hereafter NP01) used an RXTE intensive monitoring campaign on NGC 7469 to measure preliminary high temporal frequency PSDs, over the temporal frequency range $10^{-6}$ to $10^{-3}$ Hz. They did not find evidence for a break in the PSD over those temporal frequencies, but they claimed that the PSD slope flattened as photon energy increased. The PSDs we present here cover a wider temporal frequency range, and so we can revisit this claim. The rest of this paper is organized as follows. The light curve sampling and data reduction are described in Section 2. The PSD measurement and fit results are presented in Section 3. The results are discussed in Section 4, and a summary of our main conclusions is given in Section 5. Appendix A contains the results of spectral fits to summed spectra derived from the RXTE campaigns. Appendix B describes the “surrogate” Monte Carlo method of Press et al. (1992) used to derive estimates of PSD model parameter uncertainties.

2. OBSERVING STRATEGY, OBSERVATIONS, AND DATA REDUCTION

Most of the high dynamic range Seyfert PSDs have been based on similar observing strategies: regularly sampling the source
flux over multiple timescales such that the resulting set of light curves yields individual PSD segments covering complementary ranges in temporal frequency (e.g., Edelson & Nandra 1999). We follow the same strategy for NGC 7469.

We monitored the source with RXTE once every 4.27 days (64 satellite orbits) for a duration of 6.3 years, from 2003 April 8 to 2009 July 15 (Modified Julian Day [MJD] 52737–55027; observation identifiers 80152-05-*, 90154-02-*, 91138-02-*, 92108-04-*, 93144-06-*, and 94144-06*-). This sampling, which probes variability on timescales from ~a week to a few years, is henceforth called “long-term” sampling. Each observation lasted approximately 1 ks. There were six gaps due to satellite sun-angle viewing constraints in mid-February to early April of each year; each gap was ~50 days long. For this light curve, the average deviation from ephemeris was 0.23 days (5%). More intensive monitoring with RXTE was done to probe variability on timescales from a few hours to a month (“medium-term” sampling). RXTE observed NGC 7469 from 1996 June 10 at 00:55 UT until 1996 July 12 at 00:21 UT (MJD 50244.0–50276.0; observation identifiers 10315-01-*), obtaining 1 ks snapshots every orbit (6 ks) for approximately the first half of the campaign, followed by 2 ks snapshots every other orbit (13 ks). Finally, we used the light curves from two continuous XMM-Newton observations to quantify variability on timescales from ~1 hr to ~1 day.

For the RXTE sampling, we used data from the Proportional Counter Array (PCA; Swank 1998; Jahoda et al. 2006); extraction of background-subtracted light curves followed standard extraction procedures (we refer the reader to, e.g., M03 for additional details) and HEASOFT version 6.7 software. PCA STANDARD-2 data were collected from proportional counter units (PCUs) 0, 1, and 2 for the medium-term data (as PCUs 3 and 4 frequently exhibit breakdown during on-source time) and PCU 2 only for the long-term data (as by 1998, PCU 1 also began to exhibit repeated breakdown during on-source time, and PCU 0 lost its propane veto layer in 2000 May). PCU 2 is thus the best-calibrated PCU. We used standard screening criteria, including rejecting data gathered within 20 minutes of satellite passage through the South Atlantic Anomaly (SAA). The PCA background was estimated using the “L7-240” background models, appropriate for faint sources.

We extracted flux light curves; spectral fitting for each observation was done using XSPEC version 12.5.1n, assuming a Galactic column of 4.5 × 10²⁰ cm⁻² (Kalberla et al. 2005), the abundances of Wilms et al. (2000), and the cross sections of Verner et al. (1996). Response files were generated for each separate observation using PCARSP version 7.10 to account for the gradual hardening of the PCA response due to the gradual leak of xenon gas into the propane layer in each PCU. NGC 7469 contains a complex warm absorber system lying along the line of sight to the continuum source, but these absorbers do not significantly impact the spectrum above 2–3 keV (≲ 1% of continuum flux absorbed; e.g., Blustin et al. 2007) and are ignored in our modeling. Errors on each flux point were derived from the standard error of 16 s count rate light curve bins within each observation. For the long-term light curve, the total number of data points after screening was 538, with 85 (15.8%) missing due to, e.g., sun-angle constraints or screening. The medium-term light curve, binned to a timescale of 6 ks, contained 479 pts, with 73 pts (15.2%) missing due to screening or time-critical observations of other sources; most of the missing points were in the second half of the campaign as the sampling time was decreased (see NP01). We extracted light curves in the 2–10, 2–5, 5–10, and 10–20 keV bands for both the long- and medium-term timescales.

For the short-term sampling, we used public archive data from two continuous XMM-Newton observations (hereafter referred to as “short1” and “short2”). XMM-Newton observed NGC 7469 starting at 2004 November 30 at 21:12 UT (revolution 912) for a duration of 85 ks, and again starting at 2004 December 3 at 01:28 (revolution 913) for a duration of 79 ks. We downloaded the pipeline processed data (version 6.6.0) from the HEASARC archive, and used data from the European Photon Imaging Camera (EPIC) pn, which observed in Small Window mode with the medium filter. No MOS data were used as the MOS cameras were in Timing Uncompressed mode; further details of the observation can be found in Blustin et al. (2007). Using XSELECT version 2.4a, we extracted source photons from a circular region of radius 40° for each observation; backgrounds were extracted from circular regions of identical size, centered ∼3° away. We searched for background flares by inspecting the 10–13 keV pn light curves, finding none. We extracted light curves in the 2–10, 2–5, and 5–10 keV bands binned to 2000 s; variability at shorter timescales was dominated by Poisson noise (see below).

Light curve sampling parameters, including mean net source and background count rates and average observed fluxes for all light curves are listed in Table 1. Also listed in Table 1 are fractional variability amplitudes Fvar (see Vaughan et al. 2003 for definition of Fvar and its error) for each light curve.

XMM-Newton does not have coverage >10 keV, so we do not have short-term sampling in the 10–20 keV range. We investigated if the longest uninterrupted 10–20 keV PCA light curves during the longest duration observations from the 1996 intensive monitoring could be used. However, we found the resulting PSDs to be dominated by Poisson noise at temporal frequencies above ~10⁻³–³ Hz, typically, and not highly suitable for PSD analysis. We also examined 20–40 keV light curves from the High Energy X-ray Timing Experiment (HEXT) detectors aboard RXTE on all three timescales, but found the resulting PSDs to generally be dominated by Poisson noise over most temporal frequencies of interest, so we do not investigate PSDs in that energy band.

2–10 keV light curves are displayed in Figure 1 for all timescales.

3. PSD MEASUREMENT AND FIT RESULTS

The PSD measurement procedure is summarized briefly here; the reader is referred to U02 or M03 for details. Light curves were linearly interpolated across gaps. Periodograms were constructed using a discrete Fourier transform (e.g., Oppenheim & Shafer 1975) and using the normalization of Miyamoto et al. (1991) and van der Klis (1997). Following Papadakis & Lawrence (1993) and Vaughan (2005), the periodogram was logarithmically binned every factor of 1.4 in f (0.15 in the logarithm) to produce the observed PSD P(f); the two lowest temporal frequency bins were widened to accommodate three periodogram points, yielding 13, 12, 6, and 6 binned PSD points for the long, medium, short1, and short2 light curves, respectively. The observed PSDs are plotted in Figure 2 for the 2–10 keV band and in Figure 3 for the sub-bands.
shape tested, an average model since 2002 have used this procedure. For each PSD model 1992). The vast majority of AGN X-ray PSD analyses conducted 

Monte Carlo procedure outlined by U02 (based on Done et al. and to account for the effect of Poisson noise, we use the measurement distortion effects of red noise leak and aliasing, errors on each binned PSD point, to account for the PSD typically contained less than 15 periodogram points, 28 MARKOWITZ V ol. 724

The temporal frequency ranges spanned refer to the binned PSD, not the unbinned periodogram. The values of 

| Bandpass (keV) | Time | Source flux \( (10^{-11} \text{ erg cm}^{-2} \text{ s}^{-1}) \) | Mean Source \( (\text{counts s}^{-1}) \) | Mean Bkgd \( (\text{counts s}^{-1}) \) | \( F_{\text{var}} \) \( (\%) \) | Temporal Frequency Range Spanned (Hz) | \( P_{\text{sim}} \) \( (\text{Hz}^{-1}) \) |
|---------------|------|---------------------------------|--------------------------------|--------------------------------|----------------|---------------------------------|----------------|
| 2–10          | Long | 3.06 s                          | 2.94                          | 3.44                          | 22.0 ± 0.16 | 8.7 \( \times 10^{-9} \) – 1.1 \( \times 10^{-6} \) | 545 |
|               | Medium | 2.87 s                           | 2.67                          | 3.59                          | 15.93 ± 0.12 | 6.3 \( \times 10^{-7} \) – 7.2 \( \times 10^{-5} \) | 0.59 |
|               | Short1 | 2.463 s                          | 0.015                         | 8.39 ± 0.22                  | 2.0 \( \times 10^{-5} \) – 2.0 \( \times 10^{-4} \) | 0.98 |
|               | Short2 | 2.516 s                          | 0.010                         | 6.06 ± 0.22                  | 2.2 \( \times 10^{-5} \) – 2.0 \( \times 10^{-4} \) | 0.88 |
| 2–5           | Long | 1.70 s                          | 1.50                          | 1.47                          | 24.5 ± 0.22 | 8.7 \( \times 10^{-9} \) – 1.1 \( \times 10^{-6} \) | 974 |
|               | Medium | 1.56 s                           | 1.37                          | 1.62                          | 17.20 ± 0.13 | 6.3 \( \times 10^{-7} \) – 7.2 \( \times 10^{-5} \) | 1.07 |
|               | Short1 | 1.880 s                          | 0.008                         | 7.88 ± 0.24                  | 2.0 \( \times 10^{-5} \) – 2.0 \( \times 10^{-4} \) | 1.24 |
|               | Short2 | 1.904 s                          | 0.006                         | 6.02 ± 0.27                  | 2.2 \( \times 10^{-5} \) – 2.0 \( \times 10^{-4} \) | 1.17 |
| 5–10          | Long | 1.53 s                          | 1.52                          | 2.04                          | 20.5 ± 0.23 | 8.7 \( \times 10^{-9} \) – 1.1 \( \times 10^{-6} \) | 1130 |
|               | Medium | 1.28 s                           | 1.27                          | 1.89                          | 15.08 ± 0.13 | 6.3 \( \times 10^{-7} \) – 7.2 \( \times 10^{-5} \) | 1.31 |
|               | Short1 | 0.584 s                          | 0.007                         | 9.98 ± 0.46                  | 2.0 \( \times 10^{-5} \) – 2.0 \( \times 10^{-4} \) | 3.37 |
|               | Short2 | 0.613 s                          | 0.004                         | 6.20 ± 0.47                  | 2.2 \( \times 10^{-5} \) – 2.0 \( \times 10^{-4} \) | 4.09 |
| 10–20         | Long | 1.72 s                          | 0.50                          | 2.14                          | 19.92 ± 0.64 | 8.7 \( \times 10^{-9} \) – 1.1 \( \times 10^{-6} \) | 7770 |
|               | Medium | 1.70 s                           | 0.46                          | 2.38                          | 13.61 ± 0.40 | 6.3 \( \times 10^{-7} \) – 7.2 \( \times 10^{-5} \) | 9.06 |

Notes. Source fluxes are corrected for Galactic absorption. XMM-Newton fluxes were estimated from pn count rates using the HEASARC’s online W3PIMMS tool and assuming a power-law photon index of 1.7. Count rates are in units of counts s\(^{-1}\) PCU\(^{-1}\) for the long- and medium-term RXTE sampling or pn counts s\(^{-1}\) for the short-term XMM-Newton light curves. For each long-, medium-, and all short-term light curves, the sampling time \( \Delta T_{\text{samp}} \) was 368.6 ks (4.266 days), 5775 s, and 2000 s, respectively. For each of the long, medium, short1, and short2 light curves, the durations were 2290 days, 32.0 days, 85.0 ks, and 79.0 ks, respectively. The temporal frequency ranges spanned refer to the binned PSD, not the unbinned periodogram. The values of \( P_{\text{sim}} \) for the short-term light curves were measured from the high-frequency PSDs above \( 10^{-3.6} \) Hz, constructed from light curves binned to 300 s. The values of \( P_{\text{sim}} \) for the medium- and long-term light curves were estimated based on the net and background count rates (see the text for details).

The \(~5\) lowest temporal frequency bins in each individual PSD typically contained less than 15 periodogram points, precluding us from assigning normal errors. To estimate proper errors on each binned PSD point, to account for the PSD measurement distortion effects of red noise leak and aliasing, and to account for the effect of Poisson noise, we use the Monte Carlo procedure outlined by U02 (based on Done et al. 1992). The vast majority of AGN X-ray PSD analyses conducted since 2002 have used this procedure. For each PSD model shape tested, an average model \( P_{\text{sim}}(f) \) is calculated based on simulated PSDs; \( P_{\text{sim}}(f) \) accounts for the distortion effects and has errors assigned based on the rms spread of the individual simulated PSDs within each temporal frequency bin. For each model, the value of the test statistic \( \chi^2_{\text{dist}} \) between \( P_{\text{sim}}(f) \) and the observed PSD is compared to an empirical distribution of simulated \( \chi^2_{\text{dist}} \) values. The “rejection probability” \( R \) is a goodness of fit measure defined as the percentile of simulated \( \chi^2_{\text{dist}} \) values exceeded by the value of the observed \( \chi^2_{\text{dist}} \); below we list the likelihood of acceptance \( L \equiv 1 - R \).

The constant level of power due to Poisson noise is not subtracted from the data, but instead is added to each model discussed below. For the long- and medium-term PSDs, the
Figure 2. Top two panels show the observed 2–10 keV PSD, plotted in $f \times P(f)$ in panel (b) space to visually emphasize the turnover. Open squares, filled circles, gray open diamonds, and crosses denote the long, medium, short1, and short2 PSD segments, respectively. The solid lines in panels (a) and (b) denote the best-fit sharply broken power-law model folded through the sampling window (i.e., containing the distortion effects of PSD measurement and power due to Poisson noise). Panels (c), (d), and (e) show the data–model residuals to the best-fit unbroken, sharply-broken, and slowly-bending power-law models, respectively.

Power due to Poisson noise is estimated using $P_{\text{Poisson}} = 2(\mu + B)/\mu^2$, where $\mu$ and $B$ are the total net and background count rates, respectively; as these light curves are non-continuous light curves we multiply our estimate of $P_{\text{Poisson}}$ by the average value of the ratio $\Delta T_{\text{samp}}$, the average sampling time, to the average exposure time per snapshot. For each short-term PSD, we used light curves binned to 300 s to measure the PSD out to $10^{-2.8}$ Hz, well into the temporal frequency range dominated by Poisson noise (above $10^{-3.6}$ Hz, typically). A best fit to all binned PSD points in the range $10^{-3.6}$ to $10^{-2.8}$ Hz yielded the $P_{\text{Poisson}}$ values listed in Table 1.

Given that there is no overlap in time between the long-, medium-, and short-term light curves, we implicitly assume that the intrinsic variability process is only weakly non-stationary over 13 years. That is, we assume that the intrinsic PSD has remained constant in both shape and normalization for a given energy band, and the expectation value of $F_{\text{var}}$ for a given energy band also remains constant. This is a reasonable assumption: PSDs in BH XRBs tend to display significant changes in the shape or normalization of the components comprising the PSD (e.g., Lorentzians and/or power laws) on timescales of ~ a day or longer (e.g., Pottschmidt et al. 2003; Remillard & McClintock 2006); scaling with black hole mass and luminosity, AGN light curves may be expected to display strong non-stationarity, with significant changes in the observed PSD, on timescales of at least centuries to millennia.

As the long-term monitoring spans a six year duration, we can test the assumption of weak non-stationarity over this duration by splitting the long-term light curve in half (before and after MJD 53882), calculating the periodogram for each half, and using the procedure outlined by Papadakis & Lawrence (1995) to determine if the periodograms are consistent. In this method, a statistic $S$ is calculated based on the sum of the differences in power at each temporal frequency (see Appendix A of Papadakis & Lawrence 1995 for the definition of $S$); for consistent periodograms, $S$ has zero mean and a variance of 1. For the periodograms derived from each half of the 2–10, 2–5, 5–10, and 10–20 keV light curves, we calculate $S$ values of +0.08, −1.08, +0.20, and −1.10, respectively, consistent with the notion that at each energy range studied, the PSD has remained constant in shape and normalization between 2003 and 2009.

The procedure of Papadakis & Lawrence (1995) cannot be used to directly compare the long- and medium-term PSDs. However, in BH XRBs, a change in the PSD is commonly accompanied by a significant change in the energy spectrum. We thus performed fits to summed energy spectra derived from the long- and medium-term RXTE monitoring. These results are presented in Appendix A and demonstrate that the energy spectra are highly similar in form on all timescales and are thus consistent with the notion of only weak non-stationarity between 1996 and 2009.

### 3.1. PSD Model Fits

The PSDs of BH XRBs are usually of sufficient quality to model multiple components, including Lorentzians and quasi-periodic oscillations (e.g., Remillard & McClintock 2006). However, the temporal frequency resolution of AGN PSDs usually means that simple unbroken or singly broken PSD model shapes provide an adequate fit (the double-peaked profile of Ark 564 measured by McHardy et al. 2007 is a notable exception). For NGC 7469, we tested three PSD model shapes: an unbroken power law, a broken power law consisting of a sharp break, and...
Figure 3. Same as Figure 2, but for the 2–5 (left, black points), 5–10 (left, gray points), and 10–20 keV sub-bands (right).

In Appendix B, we discuss various methods for estimating the (one-dimensional) confidence ranges for each fitted model parameter for this and all subsequent PSD models. There, we discuss implementation of the “surrogate” Monte Carlo method of Press et al. (1992); the resulting confidence ranges are reported in parentheses in Table 2 for the unbroken power-law model and in Table 3 for the broken-power models discussed below.

The values of $\chi^2_{\text{dist}}/\text{dof}$ are generally poor for the 2–10, 2–5, and 5–10 keV PSDs, and the likelihoods of acceptance are low (<0.3%). The 10–20 keV PSD, lacking short-term sampling,
covers less dynamic range and the likelihood of acceptance is much higher. For the lower-energy bands, the residuals plotted in Figures 2(c) and 3(c) suggest that a more complex PSD model is required.

To test for the presence of a PSD break, we employed a power-law model with a sharp break of the form

\[ P(f) = \begin{cases} \frac{A_1}{f^\alpha_1} + A_2 \frac{f}{f_b}^{-\alpha_2}, & f \leq f_b, \\ \frac{A_1}{f^\alpha_1} \frac{f}{f_b}^{-\alpha_2}, & f > f_b, \end{cases} \]

where the normalization \( A_1 \) is the PSD amplitude at the break frequency \( f_b \), and \(-\alpha_1\) and \(-\alpha_2\) are the low- and high-frequency power-law slopes, respectively, with the constraint \( \alpha_1 < \alpha_2 \).

Break frequencies were tested in the log from \(-7.4\) to \(-4.9\) in increments of 0.1. \( \alpha_0 \) and \( \alpha_1 \) were both tested in increments of 0.1, over the ranges 1.0–3.2 and 0.0–2.0, respectively. One hundred simulated PSDs were used to determine \( \bar{P}_{\text{sim}}(f) \). The best-fit model parameters, along with likelihoods of acceptance \( L_{\text{brk}} \), are listed in Table 3. Errors listed are for one interesting parameter and were determined assuming that other parameters (except for \( A_1 \)) were fixed. Data-model residuals are plotted in Figure 2(d) for the 2–10 keV PSD and in Figure 3(d) for the sub-band PSDs. Figure 4 shows contour plots of \( \alpha_0 \) versus \( f_b \) for these three PSDs at the respective best-fit values of \( \alpha_{brk} \). Best-fit values of \( f_b \) lie near 2 \times 10^{-6} Hz for all PSDs, with best-fit values of \( \alpha_{brk} \) close to 1.0. Excluding the 10–20 keV PSD, best-fit values of \( \alpha_0 \) are found to be 1.8–1.9.

We can use the ratio of the likelihoods of acceptance \( L_{\text{brk}} / L_{\text{unbr}} \) between the broken and unbroken power-law model fits to establish that incorporating a break into the model yields a significant improvement. The likelihood ratio test \( D \equiv -2 \ln \left( L_{\text{unbr}} / L_{\text{brk}} \right) \) is distributed as a \( \chi^2 \) distribution with \( n \) degrees of freedom, where \( n \) is the difference in degrees of freedom between the two models, here equal to 2. For the 2–10, 2–5, and 5–10 keV PSDs, values of \( L_{\text{brk}} / L_{\text{unbr}} \) are in the range 17–67 and values of \( D \) span 5.7–8.4, indicating an improvement in fit with respect to the null hypothesis model (no break required) at confidence levels spanning 94.0%–98.5%. However, for the 10–20 keV PSD, the improvement in fit when adding a break to the model is not significant, with \( L_{\text{brk}} / L_{\text{unbr}} = 2.4 \) and \( D = 1.72 \), an improvement in fit over the null hypothesis model at only 58% confidence. Combined with the fact that the best-fit values of \( \alpha_{brk} \) and \( \alpha_0 \) are consistent with each other, this signifies that a break has not been robustly detected in the 10–20 keV PSD.

We also tested a more slowly bending PSD model of the form \( P(f) = \left( A_1 f^{\alpha_{brk}} \right) \left( \left( 1 + f/f_b \right)^{\alpha_{brk} - \alpha_1} \right) \) testing the same range and increments of \( f_b \), \( \alpha_{brk} \), and \( \alpha_1 \) as for the sharply broken power-law model. The best-fit model parameters and likelihoods of acceptance \( L_{\text{low}} \) are listed in Table 3, with data–model residuals plotted in Figure 2(e) for the 2–10 keV PSD and Figure 3(e) for the sub-band PSDs. Contour plots

| Bandpass (keV) | 2-10 | 5-10 | 10-20 |
|---------------|------|------|-------|
| \( f_b \) (10^{-6} Hz) | 2.0^{+0.0}_{-0.2} | 4.0^{+0.0}_{-0.3} | 2.5^{+0.1}_{-0.2} |
| \( \alpha_{brk} \) | 0.8^{+0.3}_{-0.2} | 0.9^{+0.3}_{-0.2} | 1.1^{+0.1}_{-0.2} |
| \( \alpha_1 \) | 1.8^{+0.2}_{-0.2} | 1.9^{+0.3}_{-0.2} | 1.2^{+0.3}_{-0.2} |
| \( \alpha_{brk} / \alpha_1 \) | 1.8^{+0.2}_{-0.2} | 1.9^{+0.3}_{-0.2} | 1.2^{+0.3}_{-0.2} |
| \( \chi^2_{\text{dist}} / \text{dof} \) | 3.79^{+0.0}_{-0.0} | 3.45^{+0.0}_{-0.0} | 3.11^{+0.0}_{-0.0} |
| \( L_{\text{brk}} \) (%) | 29.3^{+3.3}_{-3.3} | 39.5^{+3.3}_{-3.3} | 23.8^{+2.1}_{-2.1} |

Notes. Results from fitting the PSDs with two broken power-law models, the sharply broken (top half) and slowly bending (bottom half) models. \( \alpha_0 \) and \( \alpha_{brk} \) are the power-law slopes below and above \( f_b \), respectively. \( A_1 \) is the PSD normalization. \( L_{\text{brk}} \) is the likelihood of acceptance for this model, defined as one minus the rejection probability. An asterisk (*) denotes that the parameter uncertainty pegged at the limit of the parameter space tested. The uncertainties on \( \alpha_{brk}, f_b, \) and \( \alpha_1 \) listed next to the best-fit values for the 2–10 keV PSD are the 1σ errors estimated using the M03 method. The 68%, 90%, and 99% confidence ranges listed in parentheses were estimated using the P92 method.

| Bandpass (keV) | 2-10 | 5-10 | 10-20 |
|---------------|------|------|-------|
| \( f_b \) (10^{-6} Hz) | 1.0^{+0.6}_{-0.6} | 10.0^{+5.8}_{-5.0} | 5.0^{+5.0}_{-2.5} |
| \( \alpha_{brk} \) | 0.8^{+0.3}_{-0.2} | 0.8^{+0.3}_{-0.2} | 1.0^{+0.3}_{-0.4} |
| \( \alpha_1 \) | 1.8^{+0.2}_{-0.2} | 2.3^{+0.3}_{-0.3} | 2.1^{+0.3}_{-0.3} |
| \( \alpha_{brk} / \alpha_1 \) | 1.8^{+0.2}_{-0.2} | 2.3^{+0.3}_{-0.3} | 2.1^{+0.3}_{-0.3} |
| \( \chi^2_{\text{dist}} / \text{dof} \) | 1.38^{+0.0}_{-0.0} | 2.3^{+0.3}_{-0.3} | 2.1^{+0.3}_{-0.3} |
| \( L_{\text{brk}} \) (%) | 27.4^{+3.3}_{-3.3} | 33.8^{+3.3}_{-3.3} | 23.1^{+2.1}_{-2.1} |

Notes. Results from fitting the PSDs with two broken power-law models, the sharply broken (top half) and slowly bending (bottom half) models. \( \alpha_0 \) and \( \alpha_{brk} \) are the power-law slopes below and above \( f_b \), respectively. \( A_1 \) is the PSD normalization. \( L_{\text{brk}} \) is the likelihood of acceptance for this model, defined as one minus the rejection probability. An asterisk (*) denotes that the parameter uncertainty pegged at the limit of the parameter space tested. The uncertainties on \( \alpha_{brk}, f_b, \) and \( \alpha_1 \) listed next to the best-fit values for the 2–10 keV PSD are the 1σ errors estimated using the M03 method. The 68%, 90%, and 99% confidence ranges listed in parentheses were estimated using the P92 method.
of $\alpha_{hi}$ versus $f_b$ at the respective best-fit values of $\alpha_{lo}$ are shown in Figure 4. As with the sharply broken PSD model fits, the improvement in fit when adding a break is significant; for the 2–10, 2–5, and 5–10 keV PSDs, $L_{brk}/L_{unbr}$ spans 29–75, $D$ spans 6.8–8.6, and the improvement with respect to the unbroken power law is at confidence levels spanning 96.5%–98.6%. The evidence for a break in the 10–20 keV PSD is again unconvincing, with consistent best-fit values of $\alpha_{lo}$ and $\alpha_{hi}$ and with $L_{brk}/L_{unbr} = 2.4$ and $D = 1.77$, an improvement over the null hypothesis model at only 59% confidence. As both classes of broken power-law models fit roughly similarly and yield significant evidence for a PSD break in the 2–10, 2–5, and 5–10 keV bands, we henceforth treat both models equally in the paper.

3.2. The PSD as a Function of Photon Energy

We now examine the behavior in PSD parameters with photon energy; specifically, we focus on the apparent lack of a significant detection of a break in the 10–20 keV PSD. To quantify differences between the observed behavior of the 10–20 keV PSD and that of the PSDs at lower energies, we define $\Delta \alpha$ as the difference between the power-law slopes below and above the best-fit PSD break, using the parameter errors as listed in Table 3. For the singly broken power-law model, the average of $\Delta \alpha$ for the 2–5 and 5–10 keV PSDs is $1.0\pm0.3$, while for the 10–20 keV PSD, $\Delta \alpha$ is $0.1\pm0.2$. The difference in $\Delta \alpha$ from below 10 keV to above 10 keV suggests that the absence of a break in the 10–20 keV PSD and the appearance of a break in the <10 keV PSDs are significant at approximately the 4σ confidence level. Similarly, for the slowly bending model, the average of $\Delta \alpha$ for the 2–5 and 5–10 keV PSDs is $1.4\pm0.3$, while for the 10–20 keV PSD, $\Delta \alpha$ is $0.2\pm0.3$; the difference between these $\Delta \alpha$ values also suggests that the observed difference in PSD behavior is significant at the $\sim4\sigma$ confidence level.

Figure 5 shows an overplot of the observed PSDs in each band in data space and in “model space,” along with the best-fitting sharply broken and slowly bending power-law models. Above $\sim10^{-6}$ Hz, the 10–20 keV PSD is much flatter than the 2–10, 2–5, and 5–10 keV PSDs. Similar behavior at high temporal frequencies was reported by NP01. They constructed PSDs...
Figure 5. Best-fit sharply broken (left) and slowly bending (right) model fits to the 2–10 (black), 2–5 (red), 5–10 (green), and 10–20 keV (blue) PSDs with the PSD distortion measurement effects and power due to Poisson noise removed. The PSDs are plotted in “model space”: the solid lines denote the best-fit models, and symbols denote the differences between the observed PSD points and \( P_{\text{sim}}(f) \), plotted relative to the underlying PSD model.

(A color version of this figure is available in the online journal.)

Table 4

| Component               | Parameter                          | 1996/Medium Value | 2003–2009/Long Value |
|-------------------------|------------------------------------|-------------------|----------------------|
| Model flux              | 2–10 keV flux (erg cm\(^{-2}\) s\(^{-1}\)) | 3.24 \( \times \) 10\(^{-11}\) | 3.26 \( \times \) 10\(^{-11}\) |
| Model flux              | 10–20 keV flux (erg cm\(^{-2}\) s\(^{-1}\)) | 1.89 \( \times \) 10\(^{-11}\) | 1.89 \( \times \) 10\(^{-11}\) |
| Model flux              | 20–100 keV flux (erg cm\(^{-2}\) s\(^{-1}\)) | 4.43 \( \times \) 10\(^{-11}\) | 4.25 \( \times \) 10\(^{-11}\) |
| Galactic absorption     | \( N_{\text{H},\text{gal}} \) (cm\(^{-2}\)) | 4.45 \( \times \) 10\(^{20}\) (fixed) | 4.45 \( \times \) 10\(^{20}\) (fixed) |
| Power law               | \( \Gamma \)                        | 1.89\(^{+0.03}_{-0.01}\) | 1.95\(^{+0.05}_{-0.04}\) |
| Compton reflection      | Refl. strength \( R \)              | 0.48 \( \pm \) 0.12 | 0.68\(^{+0.21}_{-0.17}\) |
| Fe K\(\alpha\) emission line | Energy (keV)                  | 6.37\(^{+0.09}_{-0.11}\) | 6.38 \( \pm \) 0.09 |
|                        | Width \( \sigma \) (keV)            | <0.48              | <0.41                |
|                        | Intensity (photons cm\(^{-2}\) s\(^{-1}\)) | 5.2\(^{+1.0}_{-0.9}\) \( \times \) 10\(^{-5}\) | 5.2\(^{+1.5}_{-0.5}\) \( \times \) 10\(^{-5}\) |
|                        | EW (eV)                            | 150\(^{+27}_{-21}\) | 146\(^{+44}_{-21}\)   |
| Good exposure time      | PCA/HEXTE-A/HEXTE-B (ks)           | 689.4/193.5/196.5 | 375.2/49.7/108.5      |

Notes. Results from joint fits to 3–30 keV PCA and 15–100 keV HEXTE using a model consisting of a power-law component, a Compton reflection component (\textsc{pexrav}), and an Fe K\(\alpha\) emission line, all modified by the Galactic column. 20–100 keV fluxes are weighted between HEXTE clusters A and B accounting for good exposure time and the number of active detectors. For 2003–2009, the good exposure time for HEXTE-A is much less than that for HEXTE-B because HEXTE-A stopped collecting real-time background data in 2006.

from the medium-term sampling, covering temporal frequencies from \( 10^{-5.8} \) to \( 10^{-4.1} \) Hz and used uninterrupted PCA light curves with a time resolution of 16 s to probe \( 10^{-3.5} \) Hz to \( 10^{-1.5} \) Hz in the 2–4, 4–10, and 10–15 keV PSDs. NP01 fixed \( P_{\text{psd}} \) to the expected noise power level instead of leaving it as a free parameter, but this yielded a good fit at high temporal frequencies (as stated earlier, we do not use PCA to measure the 10–20 keV PSD because \( P_{\text{psd}} \) dominates). However, NP01 did not take into account PSD distortion measurement effects (the effects of which will have an energy dependence if the PSD shape itself is energy dependent). Nonetheless, the current work and NP01 both find a similar result: the 10–20 keV PSD is much flatter at higher temporal frequencies. While the 10–20 keV band lacks short-term sampling, this behavior cannot be due to an effect of sampling on the medium or long timescales, as RXTE sampled all four bands equally.

We also compared the 2–5 and 5–10 keV bands to search for any PSD evolution in energy between those two bands. We first tested for any dependence of \( f_\text{b} \) with energy assuming a universal PSD shape whose power-law slopes are independent of energy. Assuming the best-fit sharply broken model to the 2–10 keV PSD, we find consistent values for \( f_\text{b} \) for the 2–5 and 5–10 keV bands, with \( \log(f_\text{b},\text{Hz}) = -5.6 \pm 0.3 \) for each band. We find \( \log(f_\text{b},\text{Hz}) = -(6.0 \pm 0.4) \) and \( -(6.0 \pm 0.3) \) for the 2–5 and 5–10 keV bands, respectively, using the slowly bending model. We then tested for an energy dependence of \( \alpha_\text{hi} \) assuming that \( f_\text{b} \) is energy independent. For a sharply broken model with \( \log(f_\text{b},\text{Hz}) \) fixed at \( -5.7 \), the values of \( \alpha_\text{hi} \) are consistent for the 2–5 and 5–10 keV bands: 1.8 \( \pm \) 0.2 and 1.8\(^{+0.1}_{-0.2}\), respectively (1.8 \( \pm \) 0.4 for each band for the slowly bending model assuming \( \log(f_\text{b},\text{Hz}) \) is fixed at \( -6.0 \)). In conclusion, there is no significant evidence for PSD evolution in energy between the 2–5 and 5–10 keV bands.

4. DISCUSSION

4.1. NGC 7469’s Place in the \( T_\text{b}=M_{\text{BH}}-m \) Plane

The primary result of this paper is that we detect a turnover in the 2–10 keV PSD of NGC 7469 at a temporal frequency of \( f_\text{b} = 2.0^{+3.0}_{-0.8} \times 10^{-6} \) Hz or \( 1.0^{+3.0}_{-0.6} \times 10^{-6} \) Hz for the sharply or
slowly bending power-law model, respectively (using the 68% confidence limits from the P92 method); these frequencies correspond to turnover timescales 5.8 ± 3.5 days or 11.6$^{+17.5}_{-8.7}$ days, respectively.

We first determine if the measured PSD turnovers are consistent with the empirical relation between $T_b$, $M_{BH}$, and $m$ quantified by McHardy et al. (2006). As noted above, NGC 7469’s black hole mass is well studied. We use the reverberation-mapped mass from Vestergaard & Peterson (2006), who provide a “calibrated” mass estimate based on Hβ width and optical luminosity measurements of $M_{BH} = 3.34^{+0.69}_{-0.68} \times 10^7 M_\odot$. Our conclusions below do not change if we use the mass estimate from the $M_{BH}$–$\sigma_*$ relation: $\sigma_* = 131 ± 5$ km s$^{-1}$ (Nelson et al. 2004), and using the $M_{BH}$–$\sigma_*$ relation of Tremaine et al. (2002) yields $M_{BH} = 2.46^{+0.44}_{-0.34} \times 10^7 M_\odot$.

For an estimate of the bolometric luminosity, we use the value from Vasudevan & Fabian (2009) based on XMM-Newton EPIC/Optical Monitor spectral energy distribution fitting, $L_{bol} = 6 \times 10^{44}$ erg s$^{-1}$. An independent estimate of $L_{bol}$ can be based on the RXTE monitoring data: a model fit jointly to the summed spectrum RXTE PCA and HEXTE spectra constructed from all available RXTE observations of NGC 7469 (Rivers et al. 2010) indicates that the modeled average unabsorbed $2–10$ keV flux is $3.1 \times 10^{41}$ erg cm$^{-2}$ s$^{-1}$. Using a luminosity distance of 62.7 Mpc (from the NED database, using the reference frame defined by the 3K cosmic microwave background radiation), the $2–10$ keV luminosity $L_{2–10} = 1.7 \times 10^{43}$ erg s$^{-1}$. Using Marconi et al. (2004), $L_{bol} \sim 22L_{2–10} = 3.7 \times 10^{44}$ erg s$^{-1}$, very close to the value from Vasudevan & Fabian (2009), which we use below.

Using the best-fit values of the coefficients in the McHardy et al. (2006) relation, $\log(L_{bol}/T_{b,pred}) = 2.10 \log(M_{BH}/(10^7 M_\odot)) - 0.98\log(L_{bol}/(10^{44}$ erg s$^{-1}) - 2.32$, with $T_{b,pred}$ in units of days, and taking into the account the uncertainty in $M_{BH}$ listed above, we obtain $T_{b,pred} = 1.25^{+0.60}_{-0.48}$ days. Taking into account the uncertainties on the coefficients in the McHardy et al. (2006) relation, $2.10 ± 0.15, 0.98 ± 0.15$, and $2.32 ± 0.2$, respectively, we can obtain $T_{b,pred}$ values from 0.2 to 9.3 days, meaning the predicted value is consistent with both the observed sharply broken and slowly bending PSD break timescales.

4.2. Corresponding Physical Timescales

Detailed discussions on the likely physical mechanisms responsible for the PSD turnovers have appeared in many of the PSD papers cited above; here, we provide only a brief review.

We assume that the bulk of the X-ray emission originates around the black hole. The Keplerian orbital timescale $t_{orb}$ at this radius for a $3 \times 10^7 M_\odot$ black hole is 1 day. Following, e.g., Treves et al. (1988), the thermal timescale $t_{th}$ is roughly $t_{orb}/\alpha$, where $\alpha$ is the accretion disk viscosity parameter. For values of $\alpha$ of $\sim 0.1–0.2$, $t_{th}$ will be $\sim 5–10$ days, consistent with the observed PSD break timescales in NGC 7469.

One variability model that has had a measure of success in explaining the observed variability properties of Seyferts and BH XRBs, including PSD shapes, temporal frequency-dependent lags, and the rms–flux relation, involves inwardly propagating fluctuations in the local mass accretion rate (Lyubarskii 1997; Kotov et al. 2001; Arévalo & Uttley 2006). In this model, the fluctuations originate across a range of annuli in the disk and travel inward toward the central X-ray source, eventually modifying the X-ray continuum emission. The PSD breaks could correspond to the local viscosity timescale $t_{visc}$ at the outer radius of X-ray emission. $t_{visc} = t_{th}/(H/R)^2$, where $H/R$ is the ratio of the disk scale height to the radius (e.g., Treves et al. 1988). For a geometrically thin disk surrounding a $3 \times 10^7 M_\odot$ black hole, with $H/R = 1/100$, and with $\alpha$ $\sim 0.1–0.2$, $t_{visc}$ $\sim 5–10 \times 10^4$ days, far too long to be associated with the observed PSD breaks. However, for a geometrically thick disk where $H/R$ approaches 1, $t_{visc}$ approaches the thermal timescale and can match the observed PSD timescale for NGC 7469. However, Arévalo & Uttley (2006) caution that if the X-ray emission region is radially extended, one can still get a bend in the PSD due to the radial distribution in variability fluctuations, and the PSD bend does not have to correspond to a singular characteristic timescale.

4.3. Energy Dependence of the PSD

Rapid variability in the Compton hump is not a likely cause for the observed energy dependence of the PSD. The Compton hump contributes only $\sim 20%–30%$ of the total $10–20$ keV emission (see Table 3 in Appendix A), and so such extreme variability in the total $10–20$ keV emission would require the bulk of the Compton-thick reflecting material to originate within light-hours of the X-ray continuum source and to be responding to continuum variations much larger than those we observe. Nandra et al. (2000) explored variability in the absolute normalization $A_{abs}$ of the Compton hump down to timescales of 1 day, but found the observed variability in $A_{abs}$ consistent with being due to model degeneracy between $\Gamma$ and $A_{abs}$. Finally, Papadakis et al. (2001) demonstrated that the emission in 2–10 and 10–15 keV bands show high coherence over $10^{-5}$ to $10^{-4}$ Hz, suggesting a common variability mechanism for both bands.

Flattening of the power-law slope at temporal frequencies above the break with increasing energy has been claimed for some PSDs published previously, such as Mkn 766 (Vaughan & Fabian 2003; Markowitz et al. 2007), and MCG–6–30–15 (Vaughan et al. 2003); $\alpha_{2–10}$ was typically 2.5 for bandpasses centered near 0.5 keV and 2.1 near 5 keV. Moreover, modeling the broadband PSD of Ark 564 using a double-Lorentzian profile, McHardy et al. (2007) found the normalization of the higher-temporal frequency Lorentzian to increase by a factor of 1.5 from the 0.6–2.0 to the 2–10 keV bands.

Interestingly, evolution in PSD shape as a function of photon energy is not uncommonly observed in BH XRBs, and is commonly attributed to changes in the normalizations and/or peak temporal frequencies of the Lorentzian components used to fit the PSD. Energy-dependent PSD changes do not tend to be confined to any particular energy spectral state (Done & Gierliński 2005). For example, Kalemci et al. (2003, their Figure 5) studied the PSD of XTE J1650–500 during an outburst decay, modeling the PSD using a single Lorentzian profile. They noted that the 6–15 keV Lorentzian peaks at a temporal frequency a factor of 3 higher than that in the 2–6 keV band, and the normalization also increased by a factor of 3, yielding a much flatter PSD above a temporal frequency of $\sim 3$ Hz. Böck et al. (2009), modeling the 4.5–5.8 keV and 9.5–15 keV PSDs of an intermediate state of Cyg X-1 using a double-Lorentzian profile, noted that the lower temporal frequency Lorentzian decreased in normalization toward the higher-energy band, while the normalization of the other Lorentzian remained the same, yielding a flatter overall PSD above $\sim 3$ Hz in the higher-energy band. Similar changes between PSDs measured over the 2–4 and 15–71 keV bands were seen across several spectral states of Cyg X-1 by Pottscheidt et al. (2006). In addition, many quasi-periodic oscillations are detected only above a threshold.
 photon energy (e.g., Strohmayer 2001a, 2001b; Montanari et al. 2009).

Though we have modeled the PSD of NGC 7469 using a simple singly broken power law, we cannot rule out the possibility that multiple broadband components may exist in this PSD, and, speculatively, the observed PSD flattening with energy in NGC 7469 and other Seyferts could be due to changes in the relative normalizations of these components, as in Ark 564 (McHardy et al. 2007). Another possibility is that the break timescale could increase dramatically from approximately $10^{-6}$ Hz in the 2–10 keV band to $10^{-4}$ Hz or greater in the 10–20 keV band. In the context of a model incorporating inwardly propagating disk fluctuations, if the bulk of the 10–20 keV emission originates at a smaller radius compared to the 2–10 keV emission (e.g., Kotov et al. 2001), then a jump in the physical parameters of the disk with radius (e.g., if $H/R$ increases by 10) could cause vastly different values for $t_{\text{isc}}$ at different radii of the disk. A final possibility is that the break frequency in the PSD of NGC 7469 occurs at the same temporal frequency in all bands, with the power-law slope flattening with photon energy. This behavior is (at least qualitatively) consistent with the jet model of Giannios et al.

\cite{2004ApJ...601..179G} (2004; see also Kylafis et al. 2008), wherein the Comptonizing corona is identified with the base of a jet, and a PSD break corresponds to the Keplerian frequency at the radius of the jet base. PSD flattening above the break is attributed to having the temperature of jet decrease with increasing radius and having the variability of the soft input photons governed by the radius at which they are emitted from the disk.

5. CONCLUSIONS

We have presented the broadband X-ray PSD of the X-ray-typical Seyfert 1.2 NGC 7469. Preliminary PSDs for this object were published by NP01 and based on a 1996 month-long RXTE campaign which quantified variability over temporal frequencies greater than approximately $10^{-6}$ Hz. We combined these data with sampling obtained from long-term monitoring with RXTE spanning 2003–2009 and two XMM-Newton longlooks obtained in 2004. The resulting high dynamic range of the PSD, $9 \times 10^{-9}$ to $2 \times 10^{-4}$ Hz, allowed us to test simple unbroken and singly broken power-law model shapes. In the 2–10 keV PSD, we find significant evidence for a break at a best-fit temporal frequency $f_b = 1–2 \times 10^{-4}$ Hz, depending on the form of the power-law break. This corresponds to a timescale of $T_b = 6–12$ days. Given NGC 7469’s well-constrained black hole mass of $3 \times 10^7 M_\odot$, the PSD break is consistent with the empirical relation between $M_{\text{BH}}, L_{\text{bol}}$, and $T_b$ of McHardy et al. (2006) for Seyfert PSDs.

Our results include reliably derived confidence regions on the best-fit parameters on models fit to the 2–10 keV PSD. We applied the “surrogate” method of Press et al. (1992), wherein one uses Monte Carlo simulations to create a large number of synthetic data sets based on the best-fit model, and fits each data set to map out the probability distributions of model parameters. The parameter confidence ranges listed in Tables 2 and 3 thus lack the ambiguity inherent in parameter confidence regions derived from previously derived methods (Markowitz et al. 2003; Uttley et al. 2002), though we emphasize that work is still ongoing in this area (e.g., Mueller & Madejski 2009).

A break is not confirmed in the 10–20 keV band PSD, which is significantly flatter than the 2–10 keV PSD (and the 2–5 and 5–10 keV sub-bands), consistent with claims by NP01. Including the current result for NGC 7469, evidence for energy-dependent PSD evolution in Seyferts has been accumulating, and includes an energy-dependent change in the normalization of a Lorentzian component (as in Ark 564; McHardy et al. 2007), and detections of evolution in power-law slope at temporal frequencies above the break (e.g., Mkn 766: Vaughan & Fabian 2003; Markowitz et al. 2007, MCG–6–30–15: Vaughan et al. 2003). The accumulated evidence suggests that energy-dependent evolution may not be uncommon in Seyfert PSDs and may be similar to energy-dependent changes observed in the (higher-quality) PSDs of several BH XRBs, thereby corroborating other timing-based observational links between Seyferts and BH XRBs (broadband PSD shapes and the scaling of PSD break frequencies with black hole mass and accretion rate relative to Eddington; coherence; phase lags) which support the notion of similar variability mechanisms at work in both classes of objects. For further progress, the community needs a large sample of AGN PSDs observed with the same quality as XRB PSDs in order to perform the same level of detailed model fitting of Lorentzian components, power-law components, etc., and determine if and how those components vary in energy. The required PSDs would need adequately high temporal frequency resolution covering both a broad temporal frequency range and a broad photon energy range, including obtaining low Poisson noise measurements above at least 10 keV and covering temporal frequencies up to at least $10^{-3}$ Hz.

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APPENDIX A

THE X-RAY ENERGY SPECTRUM OF NGC 7469 FROM 1996 TO 2009

In this Appendix, we compare energy spectra of NGC 7469 obtained during the long-, medium-, and short-term observations to check for any evidence of significant spectral changes from 1996 to 2009 which might be associated with strong non-stationarity in the underlying variability process. As mentioned previously, in BH XRBs, significant changes in the shapes and normalizations of the broadband PSD and its components are commonly accompanied by significant changes in the energy spectrum. For example, the “low/hard” spectral state commonly seen at relatively low values of $m$ (typically 0.01–0.04 for most BH XRBs; Maccarone 2003) includes a flat energy spectrum with power-law continuum photon index $\Gamma \sim 1.6–1.8$, while in the thermally dominated “high/soft” state, a strong soft thermal component with typical $k_B T \sim 1$ keV emerges, the power-law continuum is steeper ($\Gamma \gtrsim 2$), and there is relatively weaker 2–10 keV variability (e.g., Remillard & McClintock 2006).

As pointed out by, e.g., Done & Gierliński (2005), directly analogizing Seyferts with BH XRB energy states by comparing PSD and energy spectral properties is not straightforward; complications include the effects of having vastly different disk temperatures in the two classes of objects and the fact that on timescales less than years–decades, most Seyferts are
persistent while many BH XRBs are transients. However, we can examine the energy spectra of NGC 7469 during the 1996 and 2003–2009 RXTE campaigns to check for any evidence of significant spectral changes between the two non-overlapping campaigns.

We summed all PCA and HEXTE spectra from individual observations. Extraction of spectra and PCA responses followed standard procedures, and for details, we refer the reader to Rivers et al. (2010), who present model fits to the 3–100 keV time-averaged spectrum, jointly fitting the total sum of all PCA and HEXTE data over 3–30 and 15–100 keV, respectively. They found that a good fit was obtained by using a model incorporating a power-law continuum (with no evidence for a high-energy cutoff), an Fe Kα emission line modeled with a pexrav component, and a Compton reflection hump modeled with a pexrav component in XSPEC (Magdziarz & Zdziarski 1995; inclination angle assumed to be 30°), all modified by absorption by the Galactic column. 0.5% systematics were added to the PCA spectrum.

We applied this model to the summed spectra over 1996 and 2003–2009, also jointly fitting the PCA and HEXTE data over 3–30 and 15–100 keV, respectively. Good exposure times are available for 1996/Medium-term campaign are shown in the upper panel, and the data–model residuals are shown in the lower panel.

(A color version of this figure is available in the online journal.)

Figure 6. RXTE PCA (black points) and HEXTE cluster A and B (red and green points, respectively) time-averaged spectrum and best-fit model for the 1996/medium-term campaign are shown in the upper panel, and the data–model residuals are shown in the lower panel.

(A color version of this figure is available in the online journal.)

Figure 7. Same as Figure 6, but for the time-averaged 2003–2009/long-term RXTE spectrum and best-fit model.

(A color version of this figure is available in the online journal.)

is at least in part due to the fact that Blustin et al. did not model a Compton reflection hump; omitting that component from our fits causes the best-fit value of Γ to fall to 1.80 (1996) or 1.84 (2003–2009), in better agreement with Blustin et al. (2007).

Given the good agreement in spectral fit parameters, as well as in flux and luminosity (implying that the average n̄ did not change significantly between the campaigns), we can conclude that there is no evidence for a spectral state change in NGC 7469 between 1996 and 2009, and that our assumption of weak non-stationarity is valid.

APPENDIX B
NOTES ON ESTIMATION OF MODEL PARAMETER UNCERTAINTIES

In this Appendix, we briefly discuss methods for estimating the (one-dimensional) confidence ranges for each fitted model parameter.

U02 defined a 90% confidence range for one interesting parameter using the region of parameter space for which the fit rejection probability remained below 90%; however, the sizes of the resulting confidence ranges depended critically on how close the rejection probability of a given best-fit model was to this value. For example, if a model had a rejection probability of say 88%, the 90% confidence ranges were very small (see Figure 1 of Mueller & Madejski 2009). In our Tables 2 and 3, the error on each fitted model parameter was derived using a method introduced by M03. These errors correspond to values 1σ above the rejection probability R_{nabr} for the best-fit value on a Gaussian probability distribution. For example, if the best-fit model had R_{nabr} = 95.45% (2.0σ on a Gaussian probability distribution), errors correspond to R_{nabr} = 99.73% (3.0σ). This estimate of parameter error, also used by, e.g., Markowitz (2009), was used to calculate the dependence of the size of the error range on the value of the rejection probability of the best-fit model. As noted by Mueller & Madejski (2009), however, the likelihood of acceptance (rejection probability, equivalently) is a goodness of fit measure by its definition, analogous to a p-value for $\chi^2$. For both the U02 and M03 methods, it is therefore not immediately clear to which confidence level the resulting derived error ranges really correspond.

Following Press et al. (1992, hereafter P92), we thus undertook additional Monte Carlo simulations to map out the

![Figure 6](image1.png)

![Figure 7](image2.png)
the probability distributions of each of the fitted model parameters for a given best-fit model, assuming that the model is indeed the best description of the data. As described by P92, in this “surrogate” method, the best-fit model parameters to the observed PSD (e.g., listed in Tables 2 and 3), hereafter denoted \(a_0(i)\) for a set of \(i\) parameters (\(\alpha\) for the unbroken power law; \(\alpha_{lo}, \alpha_{hi}\), and \(f_b\) for the broken power-law models) are used as “surrogates” for the true, underlying parameters \(a_{true}(i)\), and are used to generate \(j\) sets of simulated data, each with best-fit parameters \(a_{sim}(i)\). The distributions of \(a_{sim}(i)\) map out the true parameter probability distributions, provided that certain conditions (discussed below) hold.

We applied this method to the best-fit unbroken, sharply broken and slowly bending model fits to the observed 2–10 keV PSD. For each of these three models, we created \(j = 100\) simulated data sets of long, medium, short1 and short2 light curves assuming the best-fit PSD model values \(a_0(i)\) (listed in Tables 2 or 3), and found the best-fit parameter values for each by repeating the Monte Carlo fit procedure. For each data set, we stepped through the same ranges of parameters as explored above, using the same increments, and selecting different random number seeds for each trial. We thus constructed distributions of the 100 best-fit values of \(\alpha\) for the unbroken power-law model and \(\alpha_{lo}, \alpha_{hi}\), and \(f_b\) for the broken power-law models. We used the San Diego Supercomputer Center’s Triton Compute Cluster to carry out this computationally intensive process, which required a total of 9100 processor hours to fully test all three model shapes. The distributions \(a_{sim}(i)\) are plotted in Figures 8, 9, and 10 for the unbroken, sharply broken, and slowly bending power-law models, respectively.

In order for the distribution of best-fit parameters to the synthetic data sets to map out the true parameter probability distributions, the following conditions must hold: first, the distribution of \(a_{sim}(i) - a_0(i)\) is assumed to be the same as the distribution of \(a_{sim}(i) - a_{true}(i)\). The assumption is reasonable provided that, over the confidence region, the shape of the distribution does not change significantly. This likely applies to the current situation: the assumed values for \(f_b\) are not close to the edges of the temporal frequency range of the PSD. In addition, we are not assuming power-law slopes steeper than \(-1.8\); for power-law slopes steeper than \(~-2.5\) to \(-3.0\), the large amount of red noise leak increases the scatter in the amplitudes of simulated PSDs, thus potentially affecting the shape of the distribution as one considers relatively steeper PSD slopes. Second, the estimator bias, defined as the difference between \(a_0(i)\) and the average of the distribution of \(a_{sim}(i)\), must be close to \(a_0(i)\). For all power-law slopes, the estimator
For the unbroken power-law model, the central 68% of the distribution, i.e., the 68% confidence range, spanned 1.30–1.36. The 99% confidence ranges span 1.22–1.56, roughly similar to the error derived by the M03 method. We conclude that the M03 error listed in Table 2 can be regarded as a roughly 99% confidence range. We did not run the P92 method for the 2–5, 5–10, and 10–20 keV PSD as that would have required an additional 300% computing time. However, given that the M03 errors listed in Table 2 are roughly similar to that for the 2–10 keV PSD, it is reasonable to assume that the error on $\alpha$ listed in Table 2 for each of these PSDs is also a roughly 99% confidence range; for readers interested in quoting 68% or 99% confidence ranges for the parameters of the broken power-law models are listed in Table 3. Comparing these ranges with the M03 error ranges for the parameters of the sharply broken power-law model, we see that the latter tend to roughly the same size as the P92 68%–90% confidence ranges.

Future PSD work can incorporate errors on model parameters that will be estimated using a method that utilizes the Neyman construction, whose application to PSD measurement is currently under development: the reader is referred to Mueller & Madejski (2009) for details.

REFERENCES

Aref’valo, P., & Uttley, P. 2006, MNRAS, 367, 801
Blustin, A. J., et al. 2007, A&A, 466, 107
Böck, M., et al. 2009, in Proc. Conf. on X-ray Astronomy 2009: Present Status, Multi-Wavelength Approach and Future Perspectives, ed. A. Comastri, M. Cappi, & L. Angelini (Melville, NY: AIP)
Chatterjee, R., et al. 2009, ApJ, 704, 1689
Done, C., & Gierliński, M. 2005, MNRAS, 364, 208
Done, C., Madejski, G. M., Mushotzky, R. F., Turner, T. J., Koyama, K., & Kunieda, H. 1992, ApJ, 400, 138
Edelson, R., & Nandra, K. 1999, ApJ, 514, 682
Giannios, D., Kylafis, N. D., & Psaltis, D. 2004, A&A, 425, 163
Jahoda, K., Markwardt, C. B., Radueva, Y., Rots, A. H., Stark, M. J., Swank, J. H., Strohmayer, T. E., & Zhang, W. 2006, ApJS, 163, 401
Kalberla, P. M. W., Burton, B. W., Hartmann, D., Arnal, E. M., Bajaja, E., Morris, R., & Pöppel, W. G. L. 2005, A&A, 440, 775
Kalemci, E., Tomsick, J. A., Rothschild, R. E., Pottschmidt, K., Corbel, S., Wijnands, R., Miller, J. M., & Kaaret, P. 2003, ApJ, 586, 419
Kotov, O., Churazov, E., & Gilfanov, M. 2001, MNRAS, 327, 799
Kylafis, N. D., Papadakis, I. E., Reig, P., Giannios, D., & Pooley, G. G. 2006, A&A, 469, 481
Lyubarskii, Y. E. 1997, MNRAS, 292, 679
Maccarone, T. 2003, A&A, 409, 697
Magdziarz, P., & Zdziarski, A. 1995, MNRAS, 273, 837
Marconi, A., Risaliti, G., Gilli, R., Hunt, L. K., Maiolino, R., & Elvis, M. 2004, MNRAS, 351, 169
Markoff, S., Nowak, M. A., & Wilms, J. 2005, ApJ, 635, 1203
Markowitz, A. 2009, ApJ, 698, 1740
Markowitz, A., Papadakis, I., Aref’valo, P., Turner, T. J., Miller, L., & Reeves, J. N. 2007, ApJ, 656, 116
Markowitz, A., et al. 2003, ApJ, 593, 96 (M03)
McHardy, I. M., Arévalo, P., Uttley, P., Papadakis, I. E., Summons, D. P., Brinkmann, W., & Page, M. J. 2007, MNRAS, 382, 985
McHardy, I. M., Gunn, K. F., Uttley, P., & Goad, M. R. 2005, MNRAS, 359, 1469
McHardy, I. M., Koerding, E., Knigge, C., Uttley, P., & Fender, R. P. 2006, Nature, 444, 730
McHardy, I. M., Papadakis, I. E., Uttley, P., Page, M. J., & Mason, K. O. 2004, MNRAS, 348, 783
Merritt, D., & Ferrarese, L. 2001, ApJ, 547, 140
Miyamoto, S., Kimura, K., Kitamoto, S., Dotani, T., & Ebisawa, K. 1991, ApJ, 383, 784
Montanari, E., Titterchuk, L., & Frontera, F. 2009, ApJ, 692, 1597
Mueller, M., & Madejski, G. 2009, ApJ, 700, 243
Nandra, K., Le, T., George, I. M., Edelson, R. A., Mushotzky, R. F., Peterson, B. M., & Turner, T. J. 2000, ApJ, 544, 734
Nandra, K., & Papadakis, I. E. 2001, ApJ, 554, 710 (NP01)
Narayan, R., & Yi, I. 1994, ApJ, 428, L13
Nayakshin, S., Kazanas, D., & Kallman, T. R. 2000, ApJ, 537, 833
Nelson, C. H., Green, R. F., Bower, G., Gehbhardt, K., & Weistrop, D. 2004, ApJ, 615, 652
Oppenheim, A., & Schafer, R. 1975, Digital Signal Processing (Englewood Cliffs, NJ: Prentice-Hall)
Papadakis, I. E., & Lawrence, A. 1993, MNRAS, 261, 612
Papadakis, I. E., & Lawrence, A. 1995, MNRAS, 272, 161
Papadakis, I. E., Nandra, K., & Kazanas, D. 2001, ApJ, 554, 713
Peterson, B. M., et al. 2004, ApJ, 613, 682
Porquet, D., et al. 2007, A&A, 473, 67
Pottschmidt, K., Wilms, J., Nowak, M. A., Larsson, S., Zdziarski, A. A., & Pooley, G. G. 2006, Adv. Space Res., 38, 1350
Pottschmidt, K., et al. 2003, A&A, 407, 1039
Press, W. H., Teukolsky, S. A., Vetterling, W. T., & Flannery, B. P. 1992, Numerical Recipes in Fortran: The Art of Scientific Computing (2nd ed.; New York: Cambridge Univ. Press) (P92)
Remillard, R. A., & McClintock, J. E. 2006, ARA&A, 44, 49
Rivers, E., Markowitz, A., & Rothschild, R. E. 2010, ApJS, submitted
Strohmayer, T. E. 2001a, ApJ, 552, L49
Strohmayer, T. E. 2001b, ApJ, 554, L169
Swank, J. 1998, in Nuclear Phys. B (Proc. Suppl.), Vol. 69, The Active X-ray Sky: Results From BeppoSAX and Rossi-XTE, ed. L. Scarsi et al. (The Netherlands: Elsevier), 12
Tremaine, S., et al. 2002, ApJ, 574, 740
Tremes, A., Maraschi, L., & Abramowicz, M. 1988, PASP, 100, 427
Uttley, P., McHardy, I. M., & Papadakis, I. E. 2002, MNRAS, 332, 231 (U02)
van der Klis, M. 1997, in Statistical Challenges in Modern Astronomy II, ed. G. J. Babu & E. D. Feigelson (New York: Springer), 321
Vasudevan, R. V., & Fabian, A. C. 2009, MNRAS, 392, 1124
Vaughan, S. 2005, A&A, 431, 391
Vaughan, S., & Fabian, A. C. 2003, MNRAS, 341, 496
Vaughan, S., Fabian, A. C., & Nandra, K. 2003, MNRAS, 339, 1237
Verner, D. A., Ferland, G. J., Korista, K. T., & Yakovlev, D. G. 1996, ApJ, 465, 487
Vestergaard, M., & Peterson, B. M. 2006, ApJ, 641, 689
Wilms, J., Allen, A., & McCray, R. 2000, ApJ, 542, 914