Polysomnographic nighttime features of Restless Legs Syndrome: A systematic review and meta-analysis

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Background: Restless Legs Syndrome (RLS) is a common sleep disorder. Polysomnographic (PSG) studies have been used to explore the night sleep characteristics of RLS, but their relationship with RLS has not been fully analyzed and researched.

Methods: We searched the Cochrane Library electronic literature, PubMed, and EMBASE databases to identify research literature comparing the differences in polysomnography between patients with RLS and healthy controls (HCs).

Results: This review identified six studies for meta-analysis. Our research found that the rapid eye movement sleep (REM)%, sleep efficiency (SE)%, total sleep time (TST) min, and N2 were significantly decreased in patients with RLS compared with HCs, while sleep latency (SL) min, stage shifts (SS), awakenings number (AWN), wake time after sleep onset (WASO) min, N1%, rapid eye movement sleep latency (REML), and arousal index (AI) were significantly increased. Additionally, there was no significant difference among N3%, slow wave sleep (SWS)%, and apnea-hypopnea index (AHI).

Conclusion: Our findings demonstrated that architecture and sleep continuity had been disturbed in patients with RLS, which further illustrates the changes in sleep structure in patients with RLS. In addition, further attention to the underlying pathophysiological mechanisms of RLS and its association with neurodegenerative diseases is needed in future studies.

Keywords
Restless Legs Syndrome, Polysomnographic, meta-analysis, pathophysiology, sleep

Introduction

Restless Legs Syndrome (RLS), also known as Willis-Ekbom disease (1), is a common neurological sensory-motor disorder, mainly characterized by strong discomfort of the lower limbs at night or at rest and the irresistible desire to move the legs (2). The prevalence of RLS is ~0.1–15% (3). Currently, RLS can be divided into idiopathic and secondary forms (4). Alternately, secondary RLS may be related to other medical...
causes, such as anemia, pregnancy chronic renal failure, and iron deficiency (5). Previous genome-wide association studies had demonstrated that MEIS1, LBXCOR1, and BTBD9 are the RLS-predisposing genes (6–8), which increase the risk of RLS (9). Although the exact pathophysiological mechanisms of RLS are unknown, there is growing research evidence that dopaminergic neurons and iron deficiency are linked with the pathogenesis of RLS (3).

Polysomnography (PSG) can distinguish between non-rapid eye movement (NREM) sleep and rapid eye movement (REM) sleep, as well as stages N1–N3 of NREM sleep, which are recognized as the gold standard for assessing the sleep amount and quality of objective sleep (10). Sleep interruptions are also useful for understanding the neurobiology of neurodegeneration, according to PSG (11). For the diagnosis of RLS, PSG gives objective evidence (10, 12). A previous imaging investigation had demonstrated that there was a negative correlation between the fractional anisotropy (FA) values in the left corticospinal tract and the number of movement arousal index (MAI) in patients with RLS, implying clinical significance (13). Although previous studies have attempted to investigate changes in the nighttime features of patients with RLS, there may be heterogeneity between studies involving clinical variables (such as disease duration, medication status, and disease type), demographic characteristics (such as sex and age), and research methods (such as PSG scoring methods and adaptation nights), so the exact differences in sleep features between healthy controls (HCs) and patients with RLS have not been fully established. In various neurological illnesses, a meta-analysis for PSG parameters has been used (10, 11, 14). Furthermore, using meta-analysis to examine changes in PSG parameters in patients with RLS can not only overcome the limitations of a single study with small sample size, but it can also be used to assess the potential influencing factors that influence changes in nighttime features via a subgroup analysis.

To our knowledge, there has been no meta-analysis investigation into PSG-measured sleep in RLS. To close this gap, we extensively evaluated prior case-control studies and, where applicable, performed meta-analytic approaches to determine the pooled effect sizes for variations in PSG variables between patients with RLS and HCs. We also looked into factors that could contribute to study heterogeneity.

Methods

Information sources and search

Before 10 September 2021, two investigators (Z.-Z.Y. and P.-F.X.) conducted a systematic review of English-language and peer-reviewed articles from PubMed, EMBASE, and the Cochrane Library databases, with no restrictions on publication type or language. Polysomnography OR sleep architecture OR sleep stages OR sleep recordings AND Restless Leg Syndrome were among the database search criteria. On March 2, 2022, we re-ran the literature search using the same search strategy to find newly published articles. To discover qualifying articles, the full texts of possibly relevant articles were retrieved. In addition, we searched references of primary research and reviewed articles accordingly to prevent omissions. Any discrepancies were passed on to a third reviewer. The International Prospective Register of Systematic Reviews (PROSPERO) was used to register our meta-analysis (ID: CRD42021254140) (15).

Inclusion criteria

In our study, mainly case-control studies were included to measure the nocturnal PSG differences between HCs and patients with RLS. The inclusion criteria for studies included in our final meta-analysis were as follows: (1) the International Classification of Sleep Disorders (16) or the International RLS Study Group (IRLSSG) (17) was used to define whether the patients meet the diagnostic criteria of RLS; (2) including the HCs group; (3) these studies provide data on night-related sleep parameters in subjects, which were obtained by PSG measurements; (4) studies published in English-language and peer-reviewed journals; and (5) observational or cohort studies.

Exclusion criteria

The exclusion criteria were as follows: (1) guidelines, case series, case reports, reviews, and letters; (2) animal studies; (3) research not associated with RLS; (4) RLS secondary to Parkinson’s disease (PD), peripheral neuropathy, pregnancy, iron deficiency, renal failure, chronic kidney disease, and drug-induced factors; (5) combined with other sleep disorders (for instance obstructive sleep apnea syndrome (OSAS) and rapid eye movement sleep behavior disorder (RBD), or narcolepsy); (6) no studies have reported the nocturnal PSG data of patients with RLS and HCs; and (7) the sleep parameter data format reported in the study cannot be converted into averages and standard deviations (SDs).

Data collection process

Data were extracted independently using a pre-designed form and completed by two investigators. Any disagreements during the data extraction process were addressed through adequate discussion and, if necessary, will be arbitrated through a third examiner. Extracted data were entered by one investigator and verified by two reviewers. Raw data were extracted from the results of the original study with strict quality
control, and the relevant corresponding author of the study was contacted if necessary.

In this study, the PSG variables examined include total sleep time (TST), wake time after sleep onset (WASO), sleep efficiency (SE), and percentage of N1, N2, and N3, REM sleep, and REM latency. In the American Academy of Sleep Medicine (AASM) scoring rules, N3 represents slow wave sleep (SWS) and also replaces stage 3 and stage 4 in the R&B nomenclature (18). Thus, the data for stage 3 and stage 4 in the included studies were also extracted for estimating SWS. Additional PSG variables include the periodic limb movements index (PLMI), apnea-hypopnea index (AHI), and arousal index (AI).

Demographic characteristics were recorded for each study, such as mean age, sex ratio (male/female), duration of disease (years), and body mass index (BMI) for patients with RLS and healthy subjects. In addition, we documented the diagnostic criteria for patients with RLS, medication status (i.e., medication-naive, medication-withdrawn, or medicated) in patients with RLS, and adaptation night.

Quality assessment

To identify the quality of the included studies, two investigators independently used the Newcastle-Ottawa Quality Assessment Scale (NOS) (19). The NOS consists of three components, respectively: the measurement of exposure factors, the intergroup comparability of both groups, and the selection of the study population, out of a total score of 9 (20). A total NOS score of ≥7, 4–6, and ≤3 was defined as high, medium, and lower quality studies, respectively (21). Any disagreements that arise from the analysis process were addressed through discussion. Arbitration by a third reviewer when necessary.

Statistical analysis

To enter and extract data, we used Excel software. Software version Review Manager 5.3 was used for the meta-analysis. The sample size, mean, and standard deviation (SD) of patients with RLS and HCs were entered to calculate the standardized mean difference (SMD) between each group of nighttime sleep features measured by PSG. The Q statistic and I-square ($I^2$) were calculated to test the magnitude of heterogeneity and to inform on the degree of overlap between the 95% confidence intervals (CIs) of different studies for the global effect-size estimate of each PSG variable. When the research results of each PSG variable were $p > 0.1$ and $I^2 \geq 50\%$, it can be assumed that there was homogeneity among the research results, and the fixed-effects model was used for analysis. Conversely, $p \leq 0.1$ and $I^2 \geq 50\%$ indicate the presence of heterogeneity, so the random-effects model will be used for analysis. To check for the publication bias, the Egger regression method was used; when the $p < 0.05$, bias was suggested (22). If publication bias existed, to adjust the effect sizes, Duval and Tweedie’s trim and fill test will be used (23). To identify potential sources of heterogeneity between studies, a subgroup analysis will be performed. The stability of the results was tested using sensitivity analysis. All significant values were set at $p < 0.05$ in this study.

Results

Study selection

The search identified a total of 1,297 candidate studies, of which 418 were duplicates. In addition, 583 studies were excluded by reviewing titles and abstracts. Immediately after, 296 studies were reviewed in full text. In addition, no studies eligible for inclusion were found by searching the reference list for potentially relevant studies. Finally, we determined that 27 studies existed that met the inclusion criteria, 26 of which were included in the meta-analysis (13, 24–49). Since one of the studies had a score of <7 on the NOS, it was excluded (50). Figure 1 shows the process of study selection. Table 1 summarizes the characteristics of the included studies.

Description of the included studies

In our study, a total of 1,101 patients with RLS and 711 HCs were included. Of the 26 studies included, the number of participate included ranged from 8 (4 patients with RLS and 4 HCs) to 175 (102 patients with RLS and 73 HCs). The mean age range of the patients with RLS and HCs included in the study was 8.4–64 years (reported in 26 studies). The percentage of female patients across studies ranged from 0 to 100% (reported in 23 studies). Of the studies that included PSG, 25 of them excluded patients with RLS who were treated with medications whose use affected sleep quality (i.e., antidepressants, benzodiazepines, and modafinil), and two studies (32, 33) did not report the medication status of patients with RLS. All the included studies performed PSG in the sleep lab. The diagnostic criteria, medication status, and PSG nights used for the analysis are presented in Table 1.

Research quality evaluation results

To assess the quality of the included studies, we used the Newcastle-Ottawa Scale (NOS). Based on the final total score for each study, the high quality of the literature can be inferred in
our study. The details of the quality assessment for each study are presented in Table 2.

Results of individual studies

Not all of the 15 sleep parameters were measured in all of the 26 case-control studies included in our study (SWS%, REM%, REML, SE%, SL, TST, SS, AWN, WASO, N1%, N2%, N3%, AHI, AI, and PLMI). The sleep parameters tested ranged from a maximum of 10 to at least 3 sleep parameters recorded. Pooled effect sizes for a specific sleep parameter were calculated by up to 26 studies (i.e., TST) and at least 3 studies (i.e., REMI). The sleep parameters included in the different studies all have the same definition and units.
| Number | Study selection          | Country     | Type of study      | Health status     | Sample size | age         | Sex ratio (male/female) | Medication | PSG night | Diagnostic criteria |
|--------|--------------------------|-------------|--------------------|------------------|-------------|-------------|------------------------|------------|-----------|---------------------|
| 1      | Martin Michau (22)       | Canada      | Case-control study | General population | 100         | 50          | 48.8 ± 11.5            | 60/40      | 29/21     | NO                  | 1 IRLSSG     |
| 2      | Diego                    | Spain       | Case-control study | General population | 12          | 12          | 54.9 ± 11.2            | 3/7        | 3/7       | NO                  | 1 IRLSSG     |
| 3      | Plazzi G (26)            | Italy       | Case-control study | General population | 17          | 17          | 42.8 ± 16.95           | 6/11       | 7/10      | NO                  | 1 IRLSSG     |
| 4      | Magdolna Hornyak, (27)   | Germany     | Case-control study | General population | 20          | 20          | 50.4 ± 9.6             | 10/10      | 10/10     | NO                  | 1 IRLSSG     |
| 5      | Raffaele Ferri, (28)     | Italy       | Case-control study | General population | 20          | 12          | 47.6 ± 12.01           | 13/7       | 3/9       | NO                  | 1 IRLSSG     |
| 6      | Gwendolyn Boehm, (29)    | Germany     | Case-control study | General population | 95          | 31          | 54.6 ± 11.1            | 49/46      | 15/16     | NO                  | 1 IRLSSG     |
| 7      | Raffaele Ferri, (30)     | Italy       | Case-control study | General population | 90          | 28          | 58.2 ± 11.84           | 37/53      | 12/16     | NO                  | 1 IRLSSG     |
| 8      | Kwang Su Cha, (31)       | South Korea | Case-control study | General population | 15          | 15          | 45.73 ± 11.78          | 2/13       | 0/15      | NO                  | 1 IRLSSG     |
| 9      | Jung-Ick Byun, (32)      | South Korea | Case-control study | General population | 4           | 4           | 53 (52–57)             | 0/4        | 0/4       | NA                  | 1 IRLSSG     |
| 10     | Stefan Seidel, (33)      | Austria     | Case-control study | General population | 7           | 7           | 55.4 ± 13.1            | 2/5        | 3/4        | NA                  | 1 IRLSSG     |
| 11     | Claudia Schilling, (34)  | Germany     | Case-control study | General population | 73          | 34          | 54.8 ± 13.2            | 27/46      | 11/23     | NO                  | 1 IRLSSG     |
| 12     | De Cock VC, (35)         | France      | Case-control study | General population | 25          | 25          | 64 ± 5                 | Na         | Na         | NO                  | 1 IRLSSG     |
| 13     | Yves Dauvillier, (36)    | France      | Case-control study | General population | 108         | 45          | 61.54 (24.10; 85.03)   | 43/65      | 17/28     | NO                  | 1 IRLSSG     |
| 14     | Wetter TC, (37)          | Germany     | Case-control study | General population | 10          | 8           | 56 ± 6                 | Na         | Na         | NO                  | 1 IRLSSG     |
| 15     | Saletu B, (38)           | Austria     | Case-control study | General population | 12          | 12          | 57.2 ± 11.7            | Na         | Na         | NO                  | 1 IRLSSG     |
| 16     | Mariusz Sieminski, (47)  | Poland      | Case-control study | General population | 30          | 27          | 49.0 ± 14.9            | 12/18      | 12/15     | NO                  | 1 IRLSSG     |
| 17     | Raffaele Ferri, (40)     | Italy       | Case-control study | General population | 22          | 22          | 29.0 ± 8.62            | 13/9       | 12/10     | NO                  | 1 IRLSSG     |
| 18     | Magdolna Hornyak, (41)   | Germany     | Case-control study | General population | 45          | 45          | 47.4 ± 10.9            | 16/29      | 16/29     | NO                  | 1 IRLSSG     |
| 19     | Raffaele Ferri, (42)     | Italy       | Case-control study | General population | 34          | 13          | 57.6 ± 9.44            | 11/23      | 6/13       | NO                  | 1 IRLSSG     |
| 20     | Raffaele Ferri, (43)     | Italy       | Case-control study | General population | 27          | 32          | 8.4 ± 2.16             | 16/11      | 19/13     | NO                  | 1 IRLSSG     |
| 21     | Raffaele Ferri, (44)     | Italy       | Case-control study | General population | 27          | 14          | 53.6 ± 14.90           | 15/12      | 7/7        | NO                  | 1 IRLSSG     |
| 22     | Thireau J, (45)          | France      | Case-control study | General population | 64          | 38          | 61.5 ± 8.8             | 30/34      | 9/29       | NO                  | 1 IRLSSG     |
| 23     | Richard P. Allen, (46)   | America     | Case-control study | General population | 28          | 20          | 62.8 ± 9.9             | 13/15      | 8/12       | NO                  | 1 IRLSSG     |
| 24     | Hea Ree Park, (47)       | Korea       | Case-control study | General population | 30          | 31          | 46.3 ± 13.0            | 22/8       | 23/8      | NO                  | 1 IRLSSG     |
| 25     | Sofiene Chenini, (48)    | France      | Case-control study | General population | 102         | 73          | 58.91 (24.50; 77.22)   | 63/39      | 45/28     | NO                  | 1 IRLSSG     |
| 26     | Sofiene Chenini, (49)    | France      | Case-control study | General population | 84          | 76          | 55.14 ± 12.32          | 53/31      | 47/29     | NO                  | 1 IRLSSG     |
### TABLE 2 Quality assessment of included studies.

| Study                     | Choose | Comparability | Exposed | Total score |
|---------------------------|--------|---------------|---------|-------------|
|                           | Case definition | Representativeness | Control selection | Control definition | Control for important factor | Ascertainment of exposure | Same method of ascertainment for cases and controls | Non-response rate |
| Michau et al. (24, 52)    | 1      | 1             | 1       | 1           | 2               | 1               | 1               | 0             | 8            |
| Diego                     | 1      | 1             | 1       | 1           | 2               | 1               | 1               | 0             | 8            |
| Garcia-Borreguero et al. (25) |        |               |         |             |                 |                 |                 |               |              |
| Plazzi et al. (26)        | 1      | 1             | 1       | 1           | 2               | 1               | 0               | 0             | 7            |
| Hornyak and Feige (25)    | 1      | 1             | 1       | 1           | 2               | 1               | 0               | 0             | 7            |
| Ferri (28)                | 1      | 1             | 1       | 1           | 2               | 1               | 0               | 0             | 7            |
| Boehm et al. (29)         | 1      | 1             | 1       | 1           | 2               | 1               | 0               | 0             | 7            |
| Ferri (30)                | 1      | 1             | 1       | 1           | 2               | 1               | 1               | 0             | 7            |
| Cha et al. (31)           | 1      | 0             | 1       | 1           | 2               | 1               | 1               | 0             | 7            |
| Byun et al. (32)          | 1      | 1             | 1       | 1           | 1               | 1               | 1               | 0             | 8            |
| Seidel et al. (33)        | 1      | 1             | 1       | 1           | 1               | 1               | 1               | 0             | 8            |
| Schilling et al. (34)     | 1      | 1             | 1       | 0           | 2               | 1               | 1               | 0             | 7            |
| De Cock et al. (35)       | 1      | 1             | 1       | 1           | 2               | 1               | 1               | 0             | 8            |
| Dausillier et al. (36)    | 1      | 1             | 1       | 1           | 2               | 1               | 1               | 0             | 8            |
| Wetter et al. (37)        | 1      | 1             | 1       | 1           | 2               | 1               | 1               | 0             | 8            |
| Saleu et al. (38)         | 1      | 1             | 1       | 1           | 2               | 1               | 1               | 0             | 8            |
| Mariusz Sieminski, (47)   | 1      | 1             | 1       | 1           | 2               | 1               | 1               | 0             | 8            |
| Ferri et al. (40)         | 1      | 1             | 1       | 0           | 2               | 1               | 1               | 0             | 7            |
| Hornyak et al. (41)       | 1      | 1             | 1       | 1           | 2               | 1               | 1               | 0             | 8            |
| Ferri et al. (42)         | 1      | 1             | 1       | 1           | 2               | 1               | 0               | 0             | 7            |
| Ferri et al. (43)         | 1      | 1             | 0       | 1           | 2               | 1               | 1               | 0             | 7            |
| Ferri et al. (44)         | 1      | 1             | 1       | 1           | 2               | 1               | 1               | 0             | 7            |
| Thireau et al. (45)       | 1      | 1             | 1       | 1           | 2               | 1               | 1               | 0             | 8            |
| Richard P. Allen, (46)    | 1      | 1             | 1       | 1           | 1               | 1               | 1               | 0             | 8            |
| Park et al. (13)          | 1      | 1             | 1       | 1           | 1               | 1               | 1               | 0             | 8            |
| Chenini et al. (39, 48)   | 1      | 0             | 1       | 1           | 2               | 1               | 1               | 0             | 7            |
| Chenini et al. (49)       | 1      | 0             | 1       | 1           | 2               | 1               | 1               | 0             | 7            |
| Michau et al. (52)        | 1      | 1             | 1       | 1           | 2               | 1               | 0               | 0             | 7            |
Comparison of RLS patients and healthy controls

Some sleep parameters were transformed and unified before the statistical analysis. We conducted a meta-analysis of a total of 15 sleep parameters, TST, WASO, SE%, SL, SS, AWN, N1%, N2%, N3%, SWS%, REM%, REML, AHl, AI, and PLMI, which were included in 26, 13, 25, 16, 4, 3, 22, 23, 4, 20, 17, 3, 6, 10, and 21 studies, respectively. Regarding the macroscopic structure of sleep, meta-analysis showed that TST min [SMD = −0.37, 95% CI: (−0.56, −0.18)], SE% [SMD = −0.61, 95% CI: (−0.80, 0.41)], N2% [SMD = −0.49, 95% CI: (−0.62, 0.36)], and REM% [SMD = −0.31, 95% CI: (−0.45, −0.16)] were significantly decreased in patients with RLS compared with the HC group, and the difference was statistically significant, while the WASO min [SMD = 0.57, 95% CI: (0.11, 1.02)], SL min [SMD = 0.34, 95% CI: (0.18 0.50)], SS event/h [SMD = 0.64, 95% CI: (0.33, 0.96)], AWN event/h [SMD = 0.72, 95% CI: (0.33, 1.12)], N1% [SMD = 0.37, 95% CI: (0.15, 0.59)], REML min [SMD = 0.78, 95% CI: (0.39, 1.16)], AI event/h [SMD = 0.61, 95% CI: (0.26, 0.95)], and PLMI event/h [SMD = 1.01, 95% CI: (0.80, 1.23)] were significantly increased, and the difference was statistically significant. There was no significant difference in N3%, SWS%, and AHl index between the RLS and HC groups (p > 0.05). The pooled effect sizes for 15 sleep parameters are displayed in Table 3.

Among them, the outcome variable of WASO [SMD = 0.57, 95% CI: (0.11, 1.02)], $I^2 = 81\%$ showed great heterogeneity. Through the sensitivity analysis, it was found that Ferri's (42) research data have a greater impact on the results, resulting in greater bias, so it was deleted. The result of the meta-analysis after the deletion was SMD = 0.44, 95% CI: (0.07, 0.80), $I^2 = 64\%$. We analyzed the large heterogeneity produced by this study and considered that it was caused by poor comparability between the experimental group and the control group.

Risk of bias

Of the 15 parameters included, we showed no publication bias in any of the studies by funnel plot and Egger's test.

Discussion

Summary of findings

This is the first systematic review and meta-analysis examining the altered PSG parameters in patients with RLS to the best of our knowledge. The ability to integrate the results of individual studies with relatively small sample sizes is a significant advantage of this approach and allows for the building of a strong body of evidence to illustrate specific issues. Our research mainly found that TST min, SE%, N2, and REM% were significantly decreased in patients with RLS compared with healthy controls, while WASO min, SL min, SS, AWN, N1%, REML, AI, and PLMI were significantly increased. Additionally, there was no significant difference between N3%, SWS%, and the AHl index. The disturbance of sleep continuity and structure in patients with RLS was the main outcome we found, which further illustrates the changes in sleep structure in patients with RLS.

Sleep changes in RLS

Restless Legs Syndrome may have a serious negative impact on sleep, considering either the associated sleep cycle limb movements or the apparent sensory symptoms. Although the exact pathophysiology of RLS is largely unknown (12). Autopsy, cerebrospinal fluid, and imaging studies have demonstrated the dysregulation of circadian dynamics in patients with RLS (51). Various previous studies on RLS have revealed a decreased fluorine-18-L-dihydroxyphenylalanine (18F-dopa) uptake in the substantia nigra along with a reduction in the postsynaptic dopaminergic activity, providing insight into the role of dopaminergic dysfunction in RLS (52–54). Previous studies have confirmed the ability of the dopamine system to significantly regulate sleep-wake cycles (55). Dopaminergic dysfunction and increased PLMI have been identified by previous studies (56), which similarly affects sleep in patients with RLS. Based on the fact that PSG changes were mainly focused on quantitative sleep parameters, we further found that objective sleep parameters were disturbed in patients with RLS, which demonstrates that patients with RLS have decreased sleep amounts and poor quality of sleep.

Previous studies have widely concluded that SWS is reduced in patients with RLS (57). The disturbance of SWS can exacerbate the neurodegenerative process (58). Our study did not find a change in SWS min% levels between patients with RLS and healthy controls. We speculate that this may be related to the limited number of studies with small sample sizes, making it difficult to observe a significant association. In patients with severe RLS symptoms, increased sleep latency and decreased sleep efficiency were observed (50), which were consistent with our findings. Previous neuroimaging evidence has confirmed that thalamic abnormalities are related to RLS (31). The hyperpolarization of thalamocortical neurons provides a primary regulation for sleep spindle generation and for the reduction of sensory inputs enabling cortical sleep (46, 59). Increased thalamocortical excitation would therefore be expected to produce both the decreased stage 2 sleep and increased wake time seen in patients with RLS (46), which was consistent with our finding that N2% was significantly decreased. A previous study confirmed that the longer WASO duration was associated with symptom severity in drug-free...
TABLE 3  Summary of meta-analysis comparing patients with Restless Legs Syndrome (RLS) patients and healthy controls (HCs).

| No. of comparisons | Number of RLS samples | Number of HC samples | SMD (95% CI) | P Value | Effect model | Heterogeneity |
|-------------------|----------------------|----------------------|--------------|---------|--------------|---------------|
|                   |                      |                      |              |         |              | Q Statistic  |
|                   |                      |                      |              |         |              | df            | I² Statistic | P Value     |
| TST min           | 26                   | 1,011                | 711          | -0.37 (-0.56,0.18) | 0.0002 | RE           | 57.36        | 25          | 62%         | <0.0001     |
| WASO min          | 13                   | 277                  | 206          | 0.57 (0.11,1.02)   | 0.01   | RE           | 56.73        | 12          | 81%         | <0.0001     |
| SE%               | 26                   | 1,101                | 711          | -0.61 (-0.80,0.41) | 0.0001 | RE           | 56.81        | 25          | 61%         | <0.0001     |
| SL min            | 16                   | 454                  | 259          | 0.34 (0.18,0.50)   | 0.0001 | FE           | 13.84        | 15          | 21%         | 0.24        |
| SS (event/h)      | 4                    | 86                   | 83           | 0.64 (0.33,0.96)   | 0.0001 | FE           | 1.08         | 3           | 0%          | 0.78        |
| AWN (event/h)     | 3                    | 59                   | 51           | 0.72 (0.33,1.12)   | 0.0003 | FE           | 3.27         | 2           | 39%         | 0.20        |
| N1%               | 22                   | 805                  | 567          | 0.37 (0.15,0.59)   | 0.001  | RE           | 42.77        | 21          | 60%         | 0.0005      |
| N2%               | 23                   | 805                  | 567          | -0.49 (-0.62,0.36) | 0.0001 | FE           | 32.60        | 22          | 48%         | 0.01        |
| N3%               | 4                    | 113                  | 92           | 0.17 (-0.11,0.45)  | 0.22   | FE           | 1.27         | 3           | 0%          | 0.74        |
| SWS%              | 20                   | 582                  | 357          | -0.15 (-0.38,0.08) | 0.19   | RE           | 38.18        | 19          | 61%         | 0.0008      |
| REM%              | 17                   | 560                  | 353          | -0.31 (-0.45,0.16) | 0.0001 | FE           | 19.00        | 16          | 16%         | 0.27        |
| REML min          | 3                    | 62                   | 54           | 0.78 (0.39,1.16)   | 0.0001 | FE           | 1.44         | 2           | 0%          | 0.49        |
| AHI (event/h)     | 6                    | 138                  | 109          | 0.27 (-0.25,0.80)  | 1.02   | RE           | 7.79         | 5           | 62%         | 0.05        |
| AI (event/h)      | 10                   | 214                  | 164          | 0.61 (0.26,0.95)   | 0.0005 | RE           | 15.36        | 9           | 54%         | 0.03        |
| PLMI (event/h)    | 21                   | 789                  | 639          | 1.01 (0.80,1.23)   | 0.0001 | FE           | 5.25         | 20          | 0%          | 0.51        |

%, percentage; Q, Cochran’s Q statistic; df, degrees of freedom; WASO, wake time after sleep onset; AHI, apnea hypopnea index; AI, arousal index; PLMI, Periodic limb movement index; REM, rapid eye movement sleep; REML, rapid eye movement sleep latency; SWS, slow wave sleep; SE, sleep efficiency; SL, sleep latency; SS, stage shifts; AWN, awakenings number; SMD, standardized mean difference; TST, total sleep time; RE, random-effects model; FE, fixed-effects model.
patients with RLS (49). Compared with the control group, we found that patients with RLS had a higher WASO than the control group. REM sleep could contribute to the maintenance of neuronal homeostasis in the brain (11). A previous study has revealed that a decreased REM sleep may exacerbate neurodegeneration (60), which was consistent with our finding that REM min% was significantly lower in patients with RLS. Various previous studies have found that the prevalence rate of RLS was higher in patients with Parkinson’s disease (PD) than in the general population (61, 62). Interestingly, a recent study has also suggested that RLS could be a possible preclinical marker of PD (63). However, whether decreased REM sleep could act as a risk factor for patients with RLS developing PD should be validated in a larger population in future work.

Limitations

There are some shortcomings in our study, and the sample size should be further expanded to improve the quality of relevant studies. In addition, it is necessary to conduct statistical analysis in combination with non-parametric effect sizes, considering the non-normal distribution of some sleep parameters. Finally, patients with specific diagnoses of different subtypes of RLS should be distinguished and studied separately to explore relevant differences. It is worth noting that the difference in bedtime of each subject in the study may also be a potential source of heterogeneity between studies and affect our overall effect size. These limitations suggest that the results of the study should be interpreted with caution and point out that more research is needed.

Conclusion

Current meta-analysis shows that polysomnography can determine sleep abnormalities in patients with RLS and healthy controls. Especially in patients with RLS, WASO min, SL min, SS, AWN, N1%, REML, and AI increased. In addition, the changes in REM sleep in patients with RLS may reflect the underlying neuropathology and may be an early sign of the process of neuropathological change.

Data availability statement

The original contributions presented in the study are included in the article-supplementary material, further inquiries can be directed to the corresponding author.

Author contributions

CG and ZY: wrote first draft and statistics. PX and TZ: statistics and data collection. HZ: conceptualization, resources, and supervision. All authors approved the submitted version.

Funding

This work was supported by the Henan Medical Science and Technology Research Program (No. 202102310082) and Henan Province Medical Science and Technology Tackling Provincial Ministry Key Projects (SBGJ202102033).

Conflict of interest

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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