Information-based measure of disagreement for more than two observers: a useful tool to compare the degree of observer disagreement

Teresa Henriques1,2,3*, Luis Antunes2,3,4, João Bernardes3,5,6, Mara Matias2, Diogo Sato1 and Cristina Costa-Santos1,3

Abstract

Background: Assessment of disagreement among multiple measurements for the same subject by different observers remains an important problem in medicine. Several measures have been applied to assess observer agreement. However, problems arise when comparing the degree of observer agreement among different methods, populations or circumstances.

Methods: The recently introduced information-based measure of disagreement (IBMD) is a useful tool for comparing the degree of observer disagreement. Since the proposed IBMD assesses disagreement between two observers only, we generalized this measure to include more than two observers.

Results: Two examples (one with real data and the other with hypothetical data) were employed to illustrate the utility of the proposed measure in comparing the degree of disagreement.

Conclusion: The IBMD allows comparison of the disagreement in non-negative ratio scales across different populations and the generalization presents a solution to evaluate data with different number of observers for different cases, an important issue in real situations. A website for online calculation of IBMD and respective 95% confidence interval was additionally developed. The website is widely available to mathematicians, epidemiologists and physicians to facilitate easy application of this statistical strategy to their own data.

Background

As several measurements in clinical practice and epidemiologic research are based on observations made by health professionals, assessment of the degree of disagreement among multiple measurements for the same subjects under similar circumstances by different observers remains a significant problem in medicine. If the measurement error is assumed to be the same for every observer, independent of the magnitude of quantity, we can estimate within-subject variability for repeated measurements by the same subject with the within-subject standard deviation, and the increase in variability when different observers are applied using analysis of variance [1]. However this strategy is not appropriate for comparing the degree of observer disagreement among different populations or various methods of measurement. Bland and Altman proposed a technique to compare the agreement between two methods of medical measurement allowing multiple observations per subject [2] and later Schluter proposed a Bayesian approach [3]. However, problems arise when comparing the degree of observer disagreement between two different methods, populations or circumstances. For example, one issue is whether during visual analysis of cardiotocograms, observer disagreement in estimation of the fetal heart rate baseline in the first hour of labor is significantly different from that in the last hour of labor when different observers assess the printed one-hour cardiotocography tracings. Another issue that remains to be resolved is whether interobserver disagreement in head circumference assessment by neonologists is less than that by nurses. To answer to this question, several neonologists should evaluate the head circumference in the...
same newborns under similar circumstances, followed by calculation of the measure of interobserver agreement, and the same procedure repeated with different nurses. Subsequently, the two interobserver agreement measures should be compared to establish whether interobserver disagreement in head circumference assessment by neonatologists is less than that by nurses.

Occasionally, intraclass correlation coefficient (ICC), a measure of reliability, and not agreement [4] is frequently used to assess observer agreement in situations with multiple observers without knowing the differences between the numerous variations of the ICC [5]. Even when the appropriate form is applied to assess observer agreement, the ICC is strongly influenced by variations in the trait within the population in which it is assessed [6]. Consequently, comparison of ICC is not always possible across different populations. Moreover important inconsistencies can be found when ICC is used to assess agreement [7].

Lin’s concordance correlation coefficient (CCC) is additionally applicable to situations with multiple observers. The Pearson coefficient of correlation assesses the closeness of data to the line of best fit, modified by taking into account the distance of this line from the 45-degree line through the origin [8-13]. Lin objected to the use of ICC as a way of assessing agreement between methods of measurement, and developed the CCC. However, similarities exist between certain specifications of the ICC and CCC measures. Nickerson, C. [14] showed the asymptotic equivalence among the ICC and CCC estimators. However, Carrasco and Jover [15] demonstrated the equivalence between the CCC and a specific ICC at parameter level. Moreover, a number of limitations of ICC, such as comparability of populations and its dependence on the covariance between observers, described above, are also present in CCC [16]. Consequently, CCC and ICC to measure observer agreement from different populations are valid only when the measuring ranges are comparable [17].

The recently introduced information-based measure of disagreement (IBMD) provides a useful tool to compare the degree of observer disagreement among different methods, populations or circumstances [18]. However, the proposed measure assesses disagreement only between two observers, which presents a significant limitation in observer agreement studies. This type of study generally requires more than just two observers, which constitutes a very small sample set.

Here, we have proposed generalization of the information-based measure of disagreement for more than two observers. As sometimes in real situations some observers do not examine all the cases (missing data), our generalized IBMD is set to allow different numbers of examiners for various observations.

Methods
IBMD among more than two observers
A novel measure of disagreement, denoted ‘information-based measure of disagreement’ (IBMD), was proposed [18] on the basis of Shannon’s notion of entropy [19], described as the average amount of information contained in a variable. In this context, the sum over all logarithms of possible outcomes of the variable is a valid measure of the amount of information, or uncertainty, contained in a variable [19]. IBMD, use logarithms to measures the amount of information contained in the differences between two observations. This measure is normalized and satisfies the flowing properties: it is a metric, scaled invariant with differential weighting [18].

\[
IBMD = \frac{1}{N} \sum_{i=1}^{N} \log_2 \left( \frac{|x_{i1} - x_{i2}|}{\max(x_{i1}, x_{i2})} + 1 \right)
\]

We aim to measure the disagreement among measurements obtained by several observers, allowing different number of observations in each case. Thus, maintaining ‘N’ as the number of cases, we consider \(M_i, i = 1, \ldots, N\), as the number of observations in case i.

Therefore considering N vectors, one for each case, \((x_{11}, \ldots, x_{1M1}), \ldots, (x_{NM1}, \ldots, x_{NMN})\) with non-negative components, the generalized information-based measure of disagreement is defined as:

\[
IBMD = \frac{1}{\sum_{i=1}^{N} \sum_{j=1}^{M_i} \sum_{k=1}^{M_j} \sum_{l=1}^{M_k} \log \left( \frac{|x_{ik} - x_{jl}|}{\max(x_{ik}, x_{jl})} + 1 \right)}
\]

with the convention \[\frac{0-0}{\max(0,0)} = 0\]

This coefficient equals 0 when the observers agree or when there is no disagreement, and increases to 1 when the distance, i.e. disagreement among the observers, increases.

The standard error and confidence interval was based on the nonparametric bootstrap, by resampling the subjects/cases with replacement, in both original and generalized IBMD measures. The bootstrap uses the data from a single sample to simulate the results if new samples were repeated over and over. Bootstrap samples are created by sampling with replacement from the dataset. A good approximation of the 95% confidence interval can be obtained by computing the 2.5th and 97.5th percentiles of the bootstrap samples. Nonparametric resampling makes no assumptions concerning the distribution of the data. The algorithm for a nonparametric bootstrap is as follows [20]:
| ID | Gymnast | Rulebook | Jude 1 | Jude 2 | Jude 3 | Jude 4 | Jude 5 | Jude 6 | Jude 7 | Jude 8 |
|----|---------|----------|--------|--------|--------|--------|--------|--------|--------|--------|
| 1  | Old     | 7.10     | 7.20   | 7.00   | 7.70   | 7.10   | 7.10   | 7.00   | 7.30   |
| 2  | Old     | 9.30     | 9.70   | 8.90   | 9.60   | 8.60   | 9.50   | 9.60   | 9.70   |
| 3  | Old     | 8.90     | 8.80   | 8.10   | 9.30   | 8.50   | 8.10   | 7.60   | 8.70   |
| 4  | Old     | 8.00     | 8.10   | 7.30   | 8.70   | 7.50   | 8.70   | 7.40   | 9.50   |
| 5  | Old     | 9.10     | 9.00   | 8.20   | 9.00   | 8.20   | 9.50   | 7.80   | 8.00   |
| 6  | Old     | 9.10     | 9.20   | 8.30   | 9.10   | 9.70   | 8.90   | 9.00   | 9.20   |
| 7  | Old     | 8.90     | 9.00   | 7.70   | 9.00   | 8.00   | 9.40   | 8.00   | 7.70   |
| 8  | Old     | 8.30     | 8.70   | 8.10   | 8.90   | 7.80   | 9.20   | 7.80   | 9.30   |
| 9  | Old     | 9.30     | 9.40   | 8.20   | 9.40   | 8.80   | 9.30   | 9.20   | 9.80   |
| 10 | Old     | 9.40     | 9.80   | 9.40   | 9.70   | 9.10   | 10.00  | 9.30   | 9.60   |
| 11 | Old     | 7.70     | 8.70   | 7.60   | 9.00   | 7.70   | 8.50   | 7.70   | 7.70   |
| 12 | Old     | 9.20     | 9.70   | 8.50   | 9.60   | 8.60   | 9.90   | 9.70   | 7.40   |
| 13 | Old     | 7.40     | 7.30   | 7.10   | 7.90   | 7.10   | 7.40   | 7.00   | 7.50   |
| 14 | Old     | 8.40     | 8.90   | 7.40   | 8.60   | 7.80   | 8.10   | 7.40   | 8.90   |
| 15 | Old     | 7.40     | 7.60   | 7.10   | 8.10   | 7.20   | 7.60   | 7.10   | 8.80   |
| 16 | Old     | 9.80     | 9.90   | 9.20   | 9.80   | 9.30   | 10.00  | 9.40   | 9.60   |
| 17 | Old     | 9.60     | 9.60   | 9.50   | 9.80   | 9.10   | 9.90   | 9.40   | 9.90   |
| 18 | Old     | 9.60     | 9.80   | 9.50   | 9.80   | 8.80   | 9.90   | 9.80   | 9.20   |
| 19 | Old     | 8.50     | 9.20   | 7.80   | 9.30   | 7.90   | 9.00   | 7.70   | 9.70   |
| 20 | Old     | 7.10     | 9.50   | 8.80   | 9.40   | 8.50   | 9.60   | 7.90   | 8.50   |
| 21 | New     | 6.50     | 8.20   | 6.60   | 9.80   | 7.50   | 7.80   | 6.10   | 5.10   |
| 22 | New     | 7.00     | 9.70   | 7.60   | 9.60   | 8.30   | 6.90   | 6.70   | 8.60   |
| 23 | New     | 7.50     | 8.60   | 6.60   | 7.80   | 9.50   | 8.10   | 6.20   | 7.60   |
| 24 | New     | 8.50     | 9.00   | 8.10   | 7.00   | 8.30   | 9.40   | 6.70   | 8.00   |
| 25 | New     | 9.70     | 8.10   | 7.50   | 6.80   | 7.70   | 8.60   | 8.30   | 7.40   |
| 26 | New     | 8.00     | 9.10   | 7.40   | 9.30   | 8.30   | 9.70   | 6.00   | 9.90   |
| 27 | New     | 7.80     | 9.70   | 7.00   | 9.70   | 8.70   | 10.00  | 9.60   | 9.50   |
| 28 | New     | 9.30     | 7.90   | 8.20   | 7.80   | 6.30   | 7.40   | 6.10   | 7.20   |
| 29 | New     | 7.10     | 9.80   | 8.10   | 9.50   | 6.30   | 9.40   | 8.90   | 6.50   |
| 30 | New     | 8.90     | 9.30   | 7.90   | 6.80   | 8.20   | 9.10   | 7.90   | 6.80   |
| 31 | New     | 9.30     | 9.80   | 8.80   | 6.60   | 8.50   | 9.80   | 7.40   | 9.90   |
| 32 | New     | 7.90     | 8.20   | 6.70   | 9.40   | 7.60   | 6.10   | 7.40   | 7.10   |
| 33 | New     | 7.60     | 8.50   | 6.40   | 8.50   | 9.20   | 7.80   | 6.20   | 9.40   |
| 34 | New     | 8.60     | 8.90   | 6.50   | 9.00   | 7.70   | 9.10   | 6.50   | 7.10   |
| 35 | New     | 8.80     | 7.20   | 8.80   | 9.30   | 8.40   | 9.30   | 6.90   | 8.60   |
| 36 | New     | 8.40     | 9.30   | 7.50   | 8.70   | 7.90   | 9.60   | 7.90   | 7.90   |
| 37 | New     | 7.50     | 8.00   | 7.20   | 8.40   | 7.40   | 7.20   | 9.10   | 9.20   |
| 38 | New     | 9.70     | 9.80   | 9.50   | 9.80   | 9.00   | 9.90   | 9.40   | 9.60   |
| 39 | New     | 8.50     | 9.20   | 8.70   | 9.30   | 7.00   | 9.70   | 8.30   | 8.00   |
| 40 | New     | 7.30     | 8.70   | 7.20   | 8.10   | 7.30   | 7.30   | 7.10   | 7.20   |
1. Sample N observations randomly with replacement from the N cases to obtain a bootstrap data set.
2. Calculate the bootstrap version of IBMD.
3. Repeat steps 1 and 2 a B times to obtain an estimate of the bootstrap distribution.

For confidence intervals of 90–95 percent B should be between 1000 and 2000 [21,22]. In the results the confidence intervals were calculated with B equal to 1000.

Software for IBMD assessment

**Website**

We have developed a website to assist with the calculation of IBMD and respective 95% confidence intervals [23]. This site additionally includes computation of the intraclass correlation coefficient (ICC). Lin’s concordance correlation coefficient (CCC) and limits of agreement can also be measured when considering only two observations per subject. The website contains a description of these methods.

**PAIRSetc software**

PAIRSetc [24,25], a software that compares matched observations, provide several agreement measures, among them the ICC, the CCC and the 95% limits of agreement. This software is constantly updated with new measures introduced on scientific literature, in fact, a coefficient of individual equivalence to measure agreement, based on replicated readings proposed in 2011 by Pan et al. [26,27] and IBMD, published in 2010, were already include.

**Examples**

Two examples (one with real data and the other with hypothetical data) were employed to illustrate the utility of the IBMD in comparing the degree of disagreement.

A gymnast’s performance is evaluated by a jury according to rulebooks, which include a combination of the difficulty level, execution and artistry. Let us suppose that a new rulebook has been recently proposed and subsequently criticized. Some gymnasts and media argue that disagreement between the jury members in evaluating the gymnastics performance with the new scoring system is higher than that with the old scoring system, and therefore oppose its use. To better understand this claim, consider a random sample of eight judges evaluating a random sample of 20 gymnasts with the old rulebook, and a different random sample of 20 gymnasts with the new rulebook. In this case, each of the 40 gymnasts presented only one performance based on pre-defined compulsory exercises, and all eight judges simultaneously viewed the same performances and rated each gymnast independently, while blinded to their previous medals and performances. Both scoring systems ranged from 0 to 10. The results are presented in Table 1.

Visual analysis of the maternal heart rate during the last hour of labor can be more difficult than that during the first hour. We believe that this is a consequence of the deteriorated quality of signal and increasing irregularity of the heart rate (due to maternal stress). Accordingly, we tested this hypothesis by examining whether in visual analysis of cardiotocograms, observer disagreement in fetal heart rate baseline estimation in the first hour of labor is lower than that in the last hour of labor when different observers assess printed one-hour cardiotocography tracings. To answer this question, we evaluated the disagreement in maternal heart rate baseline estimation during the last and first hour of labor by three independent observers.

Specifically, the heart rates of 13 mothers were acquired, as secondary data collected in Nélio Mendonça Hospital, Funchal for another study, during the initial and last hour of labor, and printed. Three experienced

| Table 2 Estimation of baseline (bpm) in 26 segments of 13 traces (13 segments corresponding to the initial hour of labor and 13 to the final hour of labor) by three obstetricians |
|---------------------------------------------------------------|
| **Mother ID** | **Segment** | **Obstetrician 1** | **Obstetrician 2** | **Obstetrician 3** |
|----------------|-------------|-------------------|-------------------|-------------------|
| 1              | Initial hour | 80                | 80                | 80                |
| 2              | Initial hour | 65                | 66                | 70                |
| 3              | Initial hour | 65                | 66                | 70                |
| 4              | Initial hour | 63                | 67                | 65                |
| 5              | Initial hour | 82                | 83                | 85                |
| 6              | Initial hour | 75                | 76                | 75                |
| 7              | Initial hour | 80                | 81                | 85                |
| 8              | Initial hour | 84                | 85                | 80                |
| 9              | Initial hour | 100               | 102               | 105               |
| 10             | Initial hour | 82                | 82                | 80                |
| 11             | Initial hour | 67                | 65                | 70                |
| 12             | Initial hour | 75                | 74                | 87                |
| 13             | Initial hour | 70                | 70                | 70                |
| 1              | Last hour   | 78                | 75                | 75                |
| 2              | Last hour   | 90                | 90                | 100               |
| 3              | Last hour   | 70                | 67                | 70                |
| 4              | Last hour   | 70                | 65                | 65                |
| 5              | Last hour   | 87                | 87                | 90                |
| 6              | Last hour   | 72                | 73                | 75                |
| 7              | Last hour   | 75                | 75                | 75                |
| 8              | Last hour   | 100               | 98                | 100               |
| 9              | Last hour   | 110               | 108               | 110               |
| 10             | Last hour   | 103               | 103               | 100               |
| 11             | Last hour   | 80                | 80                | 100               |
| 12             | Last hour   | 98                | 100               | 100               |
| 13             | Last hour   | 70                | 70                | 65                |
obstetricians were asked to independently estimate the baseline of the 26 one-hour segments. Results are presented in Table 2. The study procedure was approved by the local Research Ethics Committees and followed the Helsinki declaration. All women who participate in the study gave informed consent to participate.

Results
Hypothetical data example
Using IBMD in the gymnast’s evaluation, we can compare observer disagreement and the respective confidence interval (CI) associated with each score system.

The disagreement among judges was assessed as IBMD = 0.090 (95% CI = [0.077;0.104]) considering the old rulebook and IBMD = 0.174 (95% CI = [0.154;0.192]) with new rulebook. Recalling that the value 0 of the IBMD means no disagreement (perfect agreement), these confidence intervals clearly indicate significantly higher observer disagreement in performance evaluation using the new scoring system, compared with the old system.

Real data example
The disagreement among obstetricians in baseline estimation, considering the initial hour of labor, was IBMD = 0.048 (95% CI = [0.036;0.071]), and during the last hour of labor, IBMD = 0.048 (95% CI = [0.027;0.075]). The results indicate no significant differences in the degree of disagreement among observers between the initial and last hour of labor.

Discussion
While comparison of the degree of observer disagreement is often required in clinical and epidemiologic studies, the statistical strategies for comparative analyses are not straightforward.

Intraclass correlation coefficient is several times used in this context, however sometimes without careful in choosing the correct form. Even when the correct form of ICC is used to assess agreement, its dependence on variance does not always allow the comparability of populations. Other approaches to assess observer agreement have been proposed [28–33], but comparative analysis across populations is still difficult to achieve. The recently proposed IBMD is a useful tool to compare the degree of disagreement in non-negative ratio scales [18], and its proposed generalization allowing several observers overcomes an important limitation of this measure in this type of analysis where more than two observers are required.

Conclusions
IBMD generalization provides a useful tool to compare the degree of observer disagreement among different methods, populations or circumstances and allows evaluation of data by different numbers of observers for different cases, an important feature in real situations where some data are often missing.

The free software and available website to compute generalized IBMD and respective confidence intervals facilitates the broad application of this statistical strategy.

Competing interests
There are any non-financial competing interests (political, personal, religious, ideological, academic, intellectual, commercial or any other) to declare in relation to this manuscript.

Authors’ contributions
TH, LA and CCS have made substantial contributions to article conception and design. They also have been involved in the analysis and interpretation of data and they draft the manuscript. JB have been involved in article conception and he revised the manuscript critically for important intellectual content. MM and DS were responsible for the creation of the software and also were involved in the data analysis. All authors read and approved the final manuscript.

Acknowledgements
We acknowledge Paula Pinto from Néllio Mendonça Hospital, Funchal, who allowed us to use the maternal heart rate dataset collected for her PhD studies, approved by the local Research Ethics Committees. This work was supported by the national science foundation, Fundação para a Ciência e Tecnologia, through FEDER funds though Programa Operacional Fatores de Competitividade – COMPETE through the project CS21 with the reference PTDC/IBA-CCO/099951/2008, through the project with the reference PEST-C/SAU/UI0753/2011 and through the PhD grant with the reference SFRH/BD/70858/2010.

Author details
1Health Information and Decision Sciences Department, Faculty of Medicine, University of Porto, Al. Prof. Hernâni Monteiro, 4200-319 Porto, Portugal.
2Instituto de Telecomunicações, Porto, Portugal. 3Centre for Research in Health Technologies and Information Systems (CINTESIS), Porto, Portugal.
4Computer Science Department, Faculty of Science, University of Porto, Porto, Portugal.
5Obstetrics and Gynecology Department, Faculty of Medicine, University of Porto, Porto, Portugal.
6Instituto de Engenharia Biomédica, Porto, Portugal.

Received: 16 October 2012 Accepted: 7 March 2013
Published: 22 March 2013

References
1. Bland M: An introduction to medical statistics. 3rd edition. Oxford: Oxford University Press; 2000.
2. Bland JM, Altman DG: Agreement Between Methods of Measurement with Multiple Observations Per Individual. J Biopham Stat 2007, 17(4):571–582.
3. Schuster PJ: A multivariate hierarchical Bayesian approach to measuring agreement in repeated measurement method comparison studies. BMC Med Res Methodol 2009, 9(6).
4. De Vet H: Observer Reliability and Agreement. In Encyclopedia of Biostatistics. John Wiley & Sons, Ltd, 2005.
5. Shrouf PE, Fleiss JL: Intraclass Correlations - Uses in Assessing Rater Reliability. Psychol Bull 1979, 86(2):420–428.
6. Muller R, Butts P: A critical discussion of intraclass correlation coefficients. Stat Med 1994, 13(23–24):2465–2476.
7. Costa-Santos C, Bernardes J, Ayres-de-Campos D, Costa A, Amorim-Costa C: The limits of agreement and the intraclass correlation coefficient may be inconsistent in the interpretation of agreement. J Clin Epidemiol 2011, 64(3):264–269.
8. Bornhart HK, Haber MJ, Lin LI: An overview on assessing agreement with continuous measurements. J Biopham Stat 2007, 17(4):529–569.
9. Carrasco JL, King TS, Chinchilli VM: The concordance correlation coefficient for repeated measures estimated by variance components. J Biopham Stat 2009, 19(1):100–105.
10. King TS, Chinchilli VM, Carrasco JL: A repeated measures concordance correlation coefficient. Stat Med 2007, 26(16):3095–3113.
11. Lin L, Hedayat AS, Wu W: A unified approach for assessing agreement for continuous and categorical data. J Biopharm Stat 2007, 17(4):629–652.

12. Lin LI: A concordance correlation coefficient to evaluate reproducibility. Biometrics 1989, 45(1):255–268.

13. Lin LI, Hedayat AS, Sinha B, Yang M: Statistical methods in assessing agreement: models, issues and tools. J Am Stat Assoc 2002, 97:257–270.

14. Nickerson C: A Note on "A Concordance Correlation Coefficient to Evaluate Reproducibility". Biometrics 1997, 53:1503–1507.

15. Carrasco JL, Jover L: Estimating the generalized concordance correlation coefficient through variance components. Biometrics 2003, 59(4):849–858.

16. Atkinson G, Nevill A: Comment on the Use of Concordance Correlation to Assess the Agreement between Two Variables. Biometrics 1997, 53:775–777.

17. Lin L, Chinchilli V: Rejoinder to the Letter from Atkinson and Nevill. Biometrics 1997, 53(2):777–778.

18. Costa-Santos C, Antunes L, Souto A, Bernardes J: Assessment of disagreement: a new information-based approach. Ann Epidemiol 2010, 20(7):555–561.

19. Shannon CE: The mathematical theory of communication. 1963. MD Comput 1997, 14(4):306–317.

20. Carpenter J, Bithell J: Bootstrap confidence intervals: when, which, what? A practical guide for medical statisticians. Stat Med 2000, 19(9):1141–1164.

21. Efron B, Tibshirani RJ: An Introduction to the Bootstrap (Chapman & Hall/CRC Monographs on Statistics & Applied Probability). Chapman and Hall/CRC, 1994.

22. Davison AC, Hinkley DV: Methods and their Application (Cambridge Series in Statistical and Probabilistic Mathematics). Cambridge University Press, 1997.

23. On-line calculator of IBMD: [http://disagreement.med.up.pt].

24. Abramson J: WINPEPI (PEPI-for-Windows): computer programs for epidemiologists. Epidemiologic Perspectives & Innovations 2004, 1(1):6.

25. Abramson JH: WINPEPI updated: computer programs for epidemiologists, and their teaching potential. Epidemiol Perspect Innov 2011, 8(1).

26. Pan Y, Haber M, Barnhart HX: A new permutation-based method for assessing agreement between two observers making replicated binary readings. Stat Med 2011, 30(8):839–853.

27. Pan Y, Haber M, Gao J, Barnhart HX: A new permutation-based method for assessing agreement between two observers making replicated quantitative readings. Stat Med 2012, 31(20):2249–2261.

28. Luiz RR, Costa AJ, Kale PL, Werneck GL: Assessment of agreement of a quantitative variable: a new graphical approach. J Clin Epidemiol 2003, 56(10):963–967.

29. Monti K: Folded empirical distribution function curves-mountain plots. Am Stat 1995, 33:525–527.

30. Quan H, Shih WJ: Assessing reproducibility by the within-subject coefficient of variation with random effects models. Biometrics 1996, 52(4):1195–1203.

31. Lin LI: Total deviation index for measuring individual agreement with applications in laboratory performance and bioequivalence. Stat Med 2000, 19(2):255–270.

32. Escaramis G, Ascaso C, Carrasco JL: The total deviation index estimated by tolerance intervals to evaluate the concordance of measurement devices. BMC Med Res Methodol 2010, 10:31.

33. Barnhart HX, Kosinski AS, Haber MJ: Assessing individual agreement. J Biopharm Stat 2007, 17(4):697–719.