Hormone-related diseases and prostate cancer: An English national record linkage study

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Insulin-like growth factor-I (IGF-I) and testosterone may be related to prostate cancer risk. Acromegaly is associated with clinically high IGF-I concentrations. Klinefelter’s syndrome, testicular hypofunction and hypopituitarism are associated with clinically low testosterone concentrations. We aimed to investigate whether diagnosis with these conditions was associated with subsequent prostate cancer diagnosis and mortality. We used linked English national Hospital Episode Statistics and mortality data from 1999 to 2017 to construct and follow-up cohorts of men aged ≥35 years diagnosed with (i) acromegaly (n = 2,495) and (ii) hypogonadal-associated diseases (n = 18,763); Klinefelter’s syndrome (n = 1,992), testicular hypofunction (n = 8,086) and hypopituitarism (n = 10,331). We estimated adjusted hazard ratios (HRs) and confidence intervals (CIs) for prostate cancer diagnosis and death using Cox regression in comparison with an unexposed reference cohort of 4.3 million men, who were admitted to hospital for a range of minor surgeries and conditions (n observed cases = 130,000, n prostate cancer deaths = 30,000). For men diagnosed with acromegaly, HR for prostate cancer diagnosis was 1.33 (95% CI 1.09–1.63; p = 0.005; n observed cases = 96), HR for prostate cancer death was 1.44 (95% CI 0.92–2.26; p = 0.11; n deaths = 19). Diagnosis with Klinefelter’s syndrome was associated with a lower prostate cancer risk (HR = 0.58, 95% CI 0.37–0.91; p = 0.02; n observed cases = 19) and hypopituitarism was associated with a reduction in prostate cancer death (HR = 0.53, 95% CI 0.35–0.79; p = 0.002; n deaths = 23). These results support the hypothesised roles of IGF-I and testosterone in prostate cancer development and/or progression. These findings are important because they provide insight into prostate cancer aetiology.

Introduction
Prostate cancer is the second most common cancer in men worldwide. 1 Established risk factors include age, family history, ethnicity and genetic factors. 2 Although there are large differences in global incidence rates, little is known regarding potentially modifiable risk factors. However, data from an international pooled individual-level meta-analysis of prospective studies have shown that high circulating insulin-like growth factor-I (IGF-I) concentrations are associated with an increased risk of prostate cancer 3 and that low free testosterone is associated with a reduced risk. 4

Acromegaly is an endocrine disorder (often caused by a pituitary adenoma), characterised by hypersecretion of growth hormone by the pituitary gland, generally resulting in high

Additional Supporting Information may be found in the online version of this article.

Key words: epidemiology, IGF-I, hormones, prostate cancer, record-linkage, testosterone

Abbreviations: APC: admitted patient care; CI: confidence interval; HES: hospital episode statistics; HR: hazard ratio; ICD10: International Classification of Disease, 10th revision; IGF-I: Insulin-like growth factor-I; KS: Klinefelter’s syndrome; TH: testicular hypofunction; TRT: testosterone replacement therapy

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cancer, while men diagnosed with diseases characterised by low testosterone had a lower risk of prostate cancer mortality. The findings support the role of IGF-I and testosterone in prostate cancer pathogenesis.

The data resources were obtained for permitted use in our study and ethics approval was obtained from the Central and South Bristol Multi-Centre Research Ethics Committee (04/Q2006/176) for analysis of the record-linked data.

Acromegaly and hypogonadal cohorts
The "exposure" cohorts were constructed to define the population of males diagnosed with acromegaly, KS, TH and hypopituitarism (Supporting Information Fig. S1). Cohorts of men were identified in the linked HES dataset using the International Classification of Disease, 10th revision [ICD-10] codes E22.0 (acromegaly), Q98.0-Q98.4 (KS), E29.1 (TH) and E23.0 (hypopituitarism). KS, TH and hypopituitarism are all characterised by low testosterone concentrations; therefore, these conditions were also combined into one cohort to maximise statistical power.

The inclusion criteria were men aged 35 years or more who had an admission during the study period with one of the exposure conditions, coded in any diagnosis position on the discharge record (i.e. either as a primary diagnosis or as a comorbidity during an admission for another medical or surgical procedure). Date of entry was based on the patient’s earliest known date of hospital admission for the relevant condition. All participants who had a prior prostate cancer diagnosis, who had both acromegaly and hypogonadal-associated diseases on their hospital records, or had missing covariate data (age, year of cohort entry, the region of residence, Index of Multiple Deprivation [IMD] rank; Supporting Information Fig. S1) were excluded.

It has been recommended that men who are treated for hypogonadal diseases with testosterone replacement therapy (TRT) are screened for prostate cancer,24 which may lead to detection bias. Therefore, we excluded all men who were diagnosed with prostate cancer within the first 6 months of cohort entry across all cohorts (Supporting Information Fig. S1).

Reference cohort
A reference cohort of men with no known prior record of prostate cancer was constructed by identifying men aged 35 years or more without acromegaly, KS, TH or hypopituitarism (Supporting Information Fig. S1), who were admitted to hospital for various other conditions and injuries recorded as the primary diagnosis or main operation on the hospital record: strabismus, cataract, otitis, varicose veins, haemorrhoids, upper respiratory tract infections, nasal polyps, teeth disorders, inguinal hernia, nail diseases, sebaceous cyst, internal derangement of knee, bunions, vasectomy, dislocations/sprains/strains, bruising, gall bladder disease,
appendicectomy, hip replacement, knee replacement or tonsillectomy. This diverse range of conditions was chosen on the basis that they are common and relatively minor, so that the men in the reference cohort would be broadly representative of the general population. We chose conditions/operations that were recorded as the primary diagnosis or operation in the hospital record to avoid selecting men who were coming into hospital principally for more serious or uncommon problems.

Date of entry to the reference cohort was based on the patient’s earliest known date of admission for one of these conditions. As with the exposure cohorts, those who were diagnosed with prostate cancer within the first 6 months of cohort entry were excluded (Supporting Information Fig. S1).

**Follow-up for prostate cancer diagnosis and vital status**

From cohort entry, men were followed-up through record linkage for any subsequent day-case or inpatient admission for prostate cancer (ICD-9 code 185, ICD-10 C61), recorded either as a primary diagnosis or elsewhere on the hospital record. HES APC does not contain data on deaths occurring outside hospital. To censor for death and to identify further prostate cancer cases not otherwise recorded in hospital, we linked the HES APC records were linked to mortality records obtained from the Office for National Statistics, and assigned prostate cancer case if it appeared anywhere on the death record (Supporting Information Table S1). This register was also used to identify prostate cancer mortality. Tumour subtype data were not available in these datasets; therefore prostate cancer mortality was used as a marker of tumour aggressiveness. To capture these clinically aggressive tumours, prostate cancer death was defined as the underlying cause of death only (Supporting Information Table S1).

Our analysis was not linked to the cancer registry data. To estimate the extent of under-ascertainment (i.e., the proportion of prostate cancer patients who were not identified using HES records or mortality data), the proportion of men with prostate cancer in 2009–2013 was compared with those identified using HES APC records and mortality records (Supporting Information Table S2). The extent of under-ascertainment in men with prostate cancer was assessed using the proportion of men excluded because they died within the first 6 months after cohort entry (Supporting Information Table S2).

**Table 1. Participant characteristics at cohort entry**

| Age group years, n (%) | Acromegaly (n = 2,495) | Klinefelter's Syndrome (n = 1,992) | Testicular hypofunction (n = 8,086) | Hypopituitarism (n = 10,331) | Reference cohort (n = 4,304,300) |
|------------------------|------------------------|-------------------------------|-------------------------------|---------------------------|---------------------------------|
| 35–49                  | 698 (28.0)             | 916 (46.0)                    | 2,486 (30.7)                  | 2,771 (26.8)              | 1,441,418 (33.5)               |
| 50–64                  | 1,012 (40.9)           | 691 (34.7)                    | 3,421 (42.3)                  | 3,782 (36.6)              | 1,303,855 (30.3)               |
| 65–79                  | 675 (27.1)             | 349 (17.5)                    | 1,883 (23.3)                  | 3,009 (29.1)              | 1,184,098 (27.5)               |
| 80+                    | 101 (4.1)              | 36 (1.8)                      | 296 (3.7)                     | 769 (7.4)                 | 374,929 (8.7)                  |
| Year of index admission, n (%) |                      |                               |                               |                           |                                 |
| January 1999–2003      | 1,007 (40.4)           | 513 (25.8)                    | 1,213 (15.0)                  | 2,062 (20.0)              | 1,265,482 (29.4)               |
| 2004–2008              | 590 (23.7)             | 472 (23.7)                    | 1,563 (19.3)                  | 2,265 (21.9)              | 1,208,125 (28.1)               |
| 2009–2013              | 584 (23.4)             | 617 (31.0)                    | 3,013 (37.3)                  | 3,581 (34.7)              | 1,202,353 (27.9)               |
| 2014–March 2017        | 314 (12.6)             | 390 (19.6)                    | 2,297 (28.4)                  | 2,423 (23.5)              | 628,340 (14.6)                 |
| Government office region, n (%) |                      |                               |                               |                           |                                 |
| North East             | 104 (4.2)              | 167 (8.4)                     | 509 (6.3)                     | 633 (6.1)                 | 264,976 (6.2)                  |
| North West             | 305 (12.2)             | 326 (16.4)                    | 1,709 (21.1)                  | 1,774 (17.2)              | 643,095 (14.9)                 |
| Yorkshire and Humber   | 258 (10.3)             | 203 (10.2)                    | 1,146 (14.2)                  | 1,222 (11.8)              | 452,720 (10.5)                 |
| East Midlands          | 177 (7.1)              | 145 (7.3)                     | 437 (5.4)                     | 693 (6.7)                 | 357,496 (8.3)                  |
| West Midlands          | 212 (8.5)              | 172 (8.6)                     | 717 (8.9)                     | 723 (7.0)                 | 441,696 (10.3)                 |
| East of England        | 277 (11.1)             | 236 (11.9)                    | 754 (9.3)                     | 1,023 (9.9)               | 466,580 (10.8)                 |
| London                 | 404 (16.2)             | 229 (11.5)                    | 987 (12.2)                    | 1,648 (16.0)              | 522,824 (12.2)                 |
| South East             | 464 (18.6)             | 294 (14.8)                    | 1,131 (14.0)                  | 1,533 (14.8)              | 658,130 (15.3)                 |
| South West             | 294 (11.8)             | 220 (11.0)                    | 696 (8.6)                     | 1,082 (10.5)              | 496,783 (11.5)                 |
| Index of multiple deprivation quintile, n (%) |                      |                               |                               |                           |                                 |
| 1 (highest deprivation) | 463 (18.6)             | 523 (26.3)                    | 1,831 (22.6)                  | 2,112 (20.4)              | 796,636 (18.5)                 |
| 2                      | 461 (18.5)             | 451 (22.6)                    | 1,632 (20.2)                  | 2,042 (19.8)              | 832,001 (19.3)                 |
| 3                      | 535 (21.4)             | 395 (19.8)                    | 1,567 (19.4)                  | 2,080 (20.1)              | 897,729 (20.9)                 |
| 4                      | 519 (20.8)             | 371 (18.6)                    | 1,624 (20.1)                  | 2,094 (20.3)              | 916,192 (21.3)                 |
| 5 (lowest deprivation) | 517 (20.7)             | 252 (12.7)                    | 1,432 (17.7)                  | 2,003 (19.4)              | 861,742 (20.0)                 |

Five cohorts in total are used in this analysis. All data excludes men diagnosed with prostate cancer within the first 6 months of follow-up.
cancer recorded in cancer registry data that did not have a concurrent prostate cancer HES APC and/or mortality record was calculated in the UK Biobank dataset. In this separate dataset HES APC and death data combined captured 76% of prostate cancer cases observed in the English cancer registry.

Statistical analysis
Eligible men contributed person-years from date of cohort entry until the date of first known prostate cancer diagnosis, death, or the end of the follow-up period (March 31, 2017), whichever came first.

Multivariable Cox proportional hazards regression models were used to estimate hazard ratios (HRs) of (i) incident prostate cancer and (ii) prostate cancer mortality, comparing in turn each exposed cohort with the reference cohort. HRs were adjusted for age (in 5-year groups), year of cohort entry, region of residence (nine regions), and patients’ cancer recorded in cancer registry data that did not have a concurrent prostate cancer HES APC and/or mortality record was calculated in the UK Biobank dataset. In this separate dataset HES APC and death data combined captured 76% of prostate cancer cases observed in the English cancer registry.

Figure 1. Hazard ratio* of prostate cancer incidence and mortality in men diagnosed with acromegaly, Klinefelter’s syndrome, testicular hypofunction and hypopituitarism, in comparison with the reference cohort†. ¹HRs adjusted for age (5-year groups), year of cohort entry, region of residence (nine regions), IMD rank (fifths). ²Conditions used in the reference cohort: strabismus, cataract, otitis, varicose veins, haemorrhoids, upper respiratory tract infections, nasal polyps, teeth disorders, inguinal hernia, nail diseases, sebaceous cyst, internal derangement of knee, bunions, vasectomy, dislocations/sprains/strains, bruising, gall bladder disease, appendectomy, hip replacement, knee replacement and tonsillectomy.

Figure 2. Hazard ratio* of prostate cancer incidence in men diagnosed with acromegaly in comparison with the reference cohort, stratified by age at cohort entry and time interval. ¹HRs adjusted for age (5-year groups), year of cohort entry, region of residence (nine regions), IMD rank (fifths). All figures exclude men diagnosed with prostate cancer within the first 6 months of follow-up. ²Conditions used in the reference cohort: strabismus, cataract, otitis, varicose veins, haemorrhoids, upper respiratory tract infections, nasal polyps, teeth disorders, inguinal hernia, nail diseases, sebaceous cyst, internal derangement of knee, bunions, vasectomy, dislocations/sprains/strains, bruising, gall bladder disease, appendectomy, hip replacement, knee replacement and tonsillectomy.
Hormone-associated diseases | Prostate cancer cases/ exposed cohort | Prostate cancer cases/ reference cohort | HR (95% CI)
--- | --- | --- | ---
All hypogonadal diseases | | | |
**All** | 361/18,763 | 127,299/4,304,300 | 0.94 (0.85, 1.05)
**Age at cohort entry (years)** | | | |
35–64 | 192/12,825 | 43,576/2,745,273 | 1.14 (0.99, 1.31)
65 + | 169/5,938 | 83,723/1,559,027 | 0.79 (0.68, 0.92)
**Length of follow-up** | | | |
6 months–4 years | 200/9,102 | 55,643/1,396,191 | 1.06 (0.93, 1.22)
5+ years | 161/9,661 | 71,856/2,908,109 | 0.83 (0.71, 0.97)
Klinefelter’s syndrome | | | |
**All** | 19/1,992 | 127,299/4,304,300 | 0.58 (0.37, 0.91)
**Age at cohort entry (years)** | | | |
35–64 | 11/1,807 | 43,576/2,745,273 | 0.60 (0.33, 1.09)
65 + | 8/385 | 83,723/1,559,027 | 0.56 (0.28, 1.11)
**Length of follow-up** | | | |
6 months–4 years | 8/771 | 55,643/1,396,191 | 0.58 (0.29, 1.17)
5+ years | 11/1,221 | 71,856/2,908,109 | 0.58 (0.32, 1.05)
Testicular hypofunction | | | |
**All** | 140/8,086 | 127,299/4,304,300 | 0.97 (0.82, 1.15)
**Age at cohort entry (years)** | | | |
35–64 | 88/5,007 | 43,576/2,745,273 | 1.19 (0.97, 1.47)
65 + | 52/2,179 | 83,723/1,559,027 | 0.75 (0.57, 0.98)
**Length of follow-up** | | | |
6 months–4 years | 72/3,948 | 55,643/1,396,191 | 1.02 (0.61, 1.28)
5+ years | 66/6,553 | 71,856/2,908,109 | 0.94 (0.74, 1.19)
Hypopituitarism | | | |
**All** | 227/10,331 | 127,299/4,304,300 | 0.98 (0.86, 1.12)
**Age at cohort entry (years)** | | | |
35–64 | 111/6,553 | 43,576/2,745,273 | 1.26 (1.05, 1.52)
65 + | 116/3,778 | 83,723/1,559,027 | 0.81 (0.68, 0.97)
**Length of follow-up** | | | |
6 months–4 years | 133/5,113 | 55,643/1,396,191 | 1.14 (0.96, 1.35)
5+ years | 94/5,218 | 71,856/2,908,109 | 0.83 (0.68, 1.01)

Figure 3. Hazard ratio* of prostate cancer incidence in men diagnosed with Klinefelter’s syndrome, testicular hypofunction and hypopituitarism in comparison with the reference cohort†, stratified by age at cohort entry and time interval. *HRs adjusted for age (5-year groups), year of cohort entry, region of residence (nine regions), IMD rank (fifths). All figures exclude men diagnosed with prostate cancer within the first 6 months of follow-up. †Conditions used in the reference cohort: strabismus, cataract, otitis, varicose veins, haemorrhoids, upper respiratory tract infections, nasal polyps, teeth disorders, inguinal hemia, nail diseases, sebaceous cyst, internal derangement of knee, bunions, vasectomy, dislocations/sprains/strains, bruising, gall bladder disease, appendectomy, hip replacement, knee replacement and tonsillectomy.

IMD (in fifths). Time from cohort entry was used as the underlying time variable. The analytical cohorts included 2,495 men with acromegaly, and 1,992, 8,086 and 10,331 men with KS, TH and hypopituitarism, respectively and 4.3 million in the reference cohort (Supporting Information Fig. S1).
Subgroup analyses were also conducted to assess whether associations differed by age at cohort entry (35–64; 65+ years) and by the length of follow-up (6 months–4 years; 5+ years); these categories were chosen a priori due to possible differences in prostate tumours in younger men\textsuperscript{26} and detection bias.\textsuperscript{24} For these purposes, we fitted interaction terms between the subgroup variables and hormonal disorder diagnosis, and compared the models with and without the interaction terms using likelihood ratio tests to determine statistical significance.

All tests of statistical significance were two-sided, and statistical significance was set at the 5% level. All statistical tests were undertaken using Stata/MP 14.0 (StataCorp, College Station, TX), and figures were created in R version 3.2.3.

Data availability
The data that support the findings of our study are available from NHS Digital but restrictions apply to the availability of these data, which were used under licence for the current study, and so are not publicly available. Data are however available from the authors upon reasonable request and with permission of NHS Digital (https://digital.nhs.uk/).

Results
Participant characteristics at cohort entry are displayed in Table 1. Most men entered the cohorts aged 35–64 years. Men in the KS cohort had higher levels of socioeconomic deprivation (Table 1). In the reference cohort, median follow-up time was 7.7 years (interquartile range [IQR] 3.9–12.4), 127,299 men were diagnosed with prostate cancer and 29,022 died from the disease (as the underlying cause of death; Supporting Information Fig. S1).

In the acromegaly cohort, median follow-up time was 7.7 years (IQR 3.5–13.3), during which 96 cases of prostate cancer were diagnosed and 19 men died from prostate cancer. Compared to the reference cohort, acromegaly was associated with a 33% increased risk of being diagnosed with prostate cancer (HR = 1.33, 95% CI 1.09–1.63; \( p = 0.005 \)), and a 44% increased risk of prostate cancer death (HR = 1.44, 95% CI 0.92–2.26; \( p = 0.11 \); Fig. 1).

In the hypogonadal cohorts, median follow-up times were 6.8 (IQR 3.2–11.6), 4.9 (IQR 2.3–8.9) and 5.1 years (IQR 2.4–9.2) for the KS, TH and hypopituitarism cohorts, respectively. During these times 19, 140 and 227 men, respectively were diagnosed with prostate cancer, and 2, 13 and 23 prostate cancer deaths were recorded in each cohort, respectively. When all three hypogonadal cohorts were combined, there was no evidence of an association with risk of prostate cancer diagnosis (HR = 0.94; 95% CI 0.85–1.05; \( p = 0.27 \); based on 361 incident cases), but a significant reduction in prostate cancer death (HR = 0.57, 95% CI 0.42–0.79; \( p = 0.001 \); based on 38 deaths).

Compared to the reference cohort, men diagnosed with KS had a 42% lower risk of being diagnosed with prostate cancer (HR = 0.58, 95% CI 0.37–0.91; \( p = 0.02 \); Fig. 1). There was an insufficient number of prostate cancer deaths to investigate the association of KS with prostate cancer mortality. TH was not associated with overall risk of prostate cancer diagnosis (HR = 0.97, 95% CI 0.82–1.15; \( p = 0.73 \)), but there was a suggestion of a reduction in prostate cancer mortality, based on 13 deaths (HR = 0.62, 95% CI 0.36–1.06; \( p = 0.08 \); Fig. 1).

Similarly, hypopituitarism was not associated with overall prostate cancer diagnosis (HR = 0.98, 95% CI 0.86–1.12; \( p = 0.81 \)), but was associated with a relative reduction in prostate cancer mortality, based on 23 deaths (HR = 0.53, 95% CI 0.35–0.79; \( p = 0.002 \); Fig. 1).

There was no evidence of heterogeneity in the associations of acromegaly or KS with prostate cancer diagnosis by age at cohort entry or length of follow-up (Figs. 2 and 3). There was evidence of heterogeneity in the associations with incident prostate cancer by age at cohort entry in the TH and hypopituitarism cohorts. In both cohorts, only men aged 65+ years at cohort entry had a reduced risk of prostate cancer diagnosis in comparison with the reference cohort (Fig. 3). Men aged 35–65 years at cohort entry who were diagnosed with hypopituitarism had a significantly increased risk of being diagnosed with prostate cancer.

Discussion
Our study of hormone-related diseases and prostate cancer is the largest to date and the first to find statistically significant evidence of an association between diagnosis of acromegaly and an increased risk of incident prostate cancer. Diagnosis with any hypogonadal-associated disease was associated with a reduced risk of prostate cancer mortality, and Klinefelter’s syndrome was associated with a lower risk of prostate cancer diagnosis.

Acromegaly is associated with clinically high circulating IGF-I concentrations,\textsuperscript{5–7} which may increase prostate cancer risk by increasing cell division and reducing programmed cell death.\textsuperscript{27} A recent meta-analysis of 13 studies with data on acromegaly and prostate cancer reported a 20% increased risk of prostate cancer, however, the association was not statistically significant (based on 38 prostate cancer cases).\textsuperscript{8} Individual studies included were based on between 25 and 752 men with diagnosed acromegaly, and thus were not able to explore a possible association with prostate cancer mortality. Our study, based on 2,495 men diagnosed with acromegaly and 96 prostate cancer cases, found a 33% increased risk of prostate cancer diagnosis and a 44% increased risk of prostate cancer as the underlying cause of death. The former finding was statistically significant; the latter suggests a potentially strong association but was not statistically significant, perhaps owing to the limited number of prostate cancer deaths in this cohort (\( n = 19 \)).

KS, TH and hypopituitarism are conditions associated with clinically low testosterone concentrations.\textsuperscript{9–12} Low circulating testosterone concentrations may reduce prostatic androgen receptor signalling and may therefore reduce prostate cancer risk.\textsuperscript{428–32}
In our analysis, there was no association with prostate cancer diagnosis in the combined hypogonadal cohort, although men diagnosed with KS had a 42% lower risk. This is consistent with findings from other record linkage studies that have reported a 20–30% relative risk reduction of prostate cancer incidence in men with KS. To the best of our knowledge, this is the first study to specifically investigate TH or hypopituitarism and prostate cancer risk.

While men diagnosed with TH and hypopituitarism did not have a lower risk of prostate cancer incidence, they did have a lower risk of prostate cancer mortality (in the KS cohort the number of prostate cancer deaths was too low to investigate). Given that androgen deprivation therapy is a mainstay treatment for advanced prostate cancer, these results may indicate that a low androgen environment slows inhibits prostate tumour growth, thereby reducing the risk of tumour progression and subsequent mortality, rather than affecting tumorigenesis.

In this analysis, it is possible that there may be some sources of detection bias. TRT is a common treatment for hypogonadal diseases. Although there is no clear evidence that TRT promotes prostate tumorigenesis, it is recommended that men receiving TRT are screened and regularly monitored for prostate cancer. Therefore, men receiving TRT may be more likely to attend prostate screening and be diagnosed with prostate cancer. This screening recommendation may account for the lack of an observed inverse association with incident prostate cancer, and the heterogeneity in the associations by age observed in the TH and hypopituitarism cohorts (if screening is more common in younger men with these conditions). It may also partly account for the observed reduction in prostate cancer death in these men, if one assumes that screening leads to lower mortality. Although we excluded men diagnosed with prostate cancer within the first 6 months of cohort entry to reduce detection bias, we may not have eliminated this bias. The observation that men in the KS cohort but not in the TH or hypopituitarism cohorts had a lower risk of incident prostate cancer may be due to the life-long exposure to a low testosterone environment in men with KS and/or a lower prostate cancer screening attendance in these men, possibly related to their higher levels of socioeconomic deprivation. There are currently no clinical guidelines recommending prostate cancer screening for men with acromegaly. A high proportion of prostate tumours will not become clinically relevant. By examining prostate cancer mortality as the underlying cause of death we were able to exclude asymptomatic or less aggressive tumours which are more likely to be detected due to bias.

These analyses have some other limitations. Prostate tumour stage and grade were not available, although prostate cancer mortality was used as a proxy for clinical aggressiveness. We cannot exclude the possibility that other abnormalities which are associated with acromegaly and/or hypogonadal diseases may confound the associations. Although these diseases are informative proxy measures of aberrant hormone concentrations, no data were available regarding treatment, which aims to normalise hormone concentrations. Hence, it is unclear to what extent these men have abnormal hormone concentrations over the long-term. This implies that the strengths of the associations found in our study may be underestimated. Further research could examine the possibility of shared genetic architecture between the exposure conditions and prostate cancer risk in order to further assess causality.

Although the investigated diseases were not required to be the principal diagnosis on the hospital record, we were only able to identify men who had hospital inpatient contact; therefore, we may not have identified all diagnoses of the exposure diseases or prostate cancer cases. Analysis of the separate UK Biobank dataset showed that approximately 24% of men with cancer registry data did not have a HES record of prostate cancer. Therefore, our cohort may represent men with more severe forms of both prostate cancer and the exposure diseases. Although this may be advantageous in the context of prostate cancer, this may bias our results if the extent of missing prostate cancer cases differs between the cohorts.

Strengths of this analysis include the large size of our study population and the person-based cohort design using linked hospital records from an integrated healthcare system as well as mortality records. Coverage of all hospital records and all death records in England over a nearly 20-year period enabled us to investigate associations across a variety of rare exposure diseases and prostate cancer diagnosis and mortality. We were also able to adjust for socioeconomic deprivation and explore interactions by follow-up time. Our very large and heterogeneous reference cohort of 4.3 million men was designed to represent the general population in terms of prostate cancer risk while allowing comparability with the exposed cohorts for purposes of detailed analyses.

In summary, we found that men diagnosed with acromegaly have a higher risk of prostate cancer diagnosis and possibly prostate cancer mortality. Diagnosis of a hypogonadal-associated disease was associated with a reduction in prostate cancer mortality, but there was only evidence of a relative reduction in prostate cancer diagnosis in the men who were diagnosed with KS. Overall, these results support the roles of IGF-I and testosterone in prostate cancer development and/or progression. These associations provide insight into prostate cancer aetiology.

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