A nomogram to preoperatively predict 1-year disease-specific survival in resected pancreatic cancer following neoadjuvant chemoradiation therapy

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Abstract

Objective: This study aimed to develop a nomogram to predict the 1-year survival of patients with pancreatic cancer who underwent pancreatectomy following neoadjuvant treatment with preoperatively detectable clinical parameters. Extended pancreatectomy is necessary to achieve complete tumor removal in borderline resectable and locally advanced pancreatic cancer. However, it increases postoperative morbidity and mortality rates, and should be balanced with potential benefit of long-term survival.

Methods: The medical records of patients who underwent pancreatectomy following neoadjuvant treatment from January 2005 to December 2016 at Severance Hospital were retrospectively reviewed. Medical records were collected from five international institutions from Japan and Singapore for external validation.

Results: A total of 113 patients were enrolled. The nomogram for predicting 1-year disease-specific survival was created based on 5 clinically detectable preoperative parameters as follows: age (year), symptom (no/yes), tumor size at initial diagnostic stage (cm), preoperative serum carbohydrate antigen (CA) 19-9 level after neoadjuvant treatment (<34/≥34 U/mL), and planned surgery [pancreaticoduodenectomy (PD) (pylorus-preserving PD)/distal pancreatectomy (DP)/total pancreatectomy]. Model performance was assessed for discrimination and calibration. The calibration plot showed good agreement between actual and predicted survival probabilities; the the Greenwood-Nam-D’Agostino (GND) goodness-of-fit test showed that the model was well calibrated ($\chi^2=8.24, P=0.5099$). A total of 84 patients were used for external validation. When correlating actual disease-specific survival and calculated 1-year disease-specific survival, there were significance differences according to the calculated probability of 1-year survival among the three groups ($P=0.044$).

Conclusions: The developed nomogram had quite acceptable accuracy and clinical feasibility in the decision-making process for the management of pancreatic cancer.

Keywords: Pancreatic cancer; neoadjuvant treatment; pancreatectomy; survival; nomogram

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Introduction

Pancreatic cancer is among the most fatal cancers of the gastrointestinal tract. The overall 5-year survival rate of pancreatic cancer is only approximately 6%, and it is the fourth leading cause of cancer-related deaths in the United States (1) as well as in Korea (2). Although surgical resection with clear safe margin is the only modality for curative treatment, only 20% of patients are eligible for surgical resection at the time of initial diagnosis because majority of patients are often diagnosed in the late stage and thus have poor general condition. Moreover, more than 50% of patients who undergo curative surgery develop recurrence within 1 year, and the 5-year survival rate of resected pancreatic cancer is less than 20% (1,3,4).

According to the National Comprehensive Cancer Network (NCCN) guideline (5), pancreatic cancers requiring neoadjuvant treatment are those that involve the major vascular structures and adjacent organs. When margin-negative resection is considered as the most effective treatment (6), extended pancreatectomy including combined major vascular resection and adjacent organ resection may be necessary to achieve complete tumor removal in borderline and locally advanced pancreatic cancers. However, extended pancreatic resections following combined preoperative chemotherapy and radiotherapy before surgery are known to increase the rates of postoperative morbidity and mortality.

Several randomized control studies (7-10) that compared oncologic outcomes and short-term perioperative outcomes between extended pancreatectomy and standard pancreatectomy showed that extended pancreatectomy is associated with severe complications without improving oncologic outcomes. Bhayani et al. (11) analyzed 273 extended pancreateoduodenectomies from the National Surgical Quality Improvement Project database and reported a threefold increase in major perioperative morbidity and mortality rates even after adjusting for comorbidity in such surgical strategy. Hartwig et al. (12) also recently assessed the outcome of extended pancreatectomy for borderline resectable and locally advanced pancreatic cancer and showed that extended resections were associated with increased perioperative morbidity and mortality.

Surgeons need to consider many clinical and practical factors when deciding the need for pancreatectomy, such as potential oncologic benefit of surgical approach, patients’ existing co-morbidity, expected life span of the patients, quality of life after surgery, potential postoperative complications, and surgery-related mortality. If possible, it was a good strategy to avoid unnecessary surgery in patients without surgical benefits.

In this study, we aimed to develop a nomogram to predict the 1-year survival of patients with pancreatic cancer who underwent pancreatectomy following neoadjuvant treatment with preoperatively detectable clinical parameters to ultimately provide a basis for decision-making for the appropriate management of pancreatic cancer following neoadjuvant treatment.

Materials and methods

Study population

This study protocol was approved by the Institutional Review Board of Yonsei University College of Medicine (Approval number: 4-2018-0374). The medical records of the patients who underwent pancreatectomy following neoadjuvant treatment from January 2005 to December 2016 were retrospectively reviewed. Neoadjuvant treatment was routinely performed for the patients with borderline resectable or locally advanced pancreatic cancer. Among patients in the resectable stage, neoadjuvant treatment was performed preoperatively, even if there was no major vascular involvement, but the tumor size was large and there was invasion to the surrounding organs.

In this study, we included patients who underwent neoadjuvant treatment before surgery and were able to perform curative resection in pancreatic cancer. Patients who did not attempt pancreatic resection due to distant metastasis found during surgery and patients who did not undergo curative resection due to severe major vascular invasion were excluded.

Data collection

Data on clinicopathologic variables, such as sex, age, serum carbohydrate antigen (CA) 19-9, tumor location, operation type, tumor size, tumor grade (differentiation), pathologic tumor (pT) stage, presence of lymph node metastasis, lymph node ratio (i.e., total number of metastatic lymph nodes divided by the total number of retrieved lymph node), microscopic perineural invasion and lymphovascular
invasion, and disease-specific survival were collected. Surgeons followed up all the patients, usually every 3–4 months and the follow-up continued until recurrence or death was confirmed.

**External validation**

External validation to test the accuracy of developed nomogram to predict 1-year disease-specific survival was conducted by reviewing the medical records of the patients who underwent pancreatectomy following neoadjuvant treatment for pancreatic cancer from five international institutions from Japan and Singapore.

**Statistical analysis**

When developing the prediction model, preoperatively available clinical parameters were included in the final model that achieved statistical significance in the univariate Cox model and were also considered clinically important. Based on the results of the multivariate Cox model, a nomogram was created by using the package `rms` in R software (Version 3.1.3; R Foundation for Statistical Computing, Vienna, Austria) (http://www.r-project.org/). Serum CA19-9 levels were divided into two based on a cut-off of 34 U/mL, which was based on the proposal of Contal and O’Quigley (13) as determined via log-rank test.

To quantify the discriminative capability of the final model, Harrell’s C-index was measured. Calibration was evaluated via a calibration plot, which is a graphic representation of the relationship between the observed and predicted survival, and the Greenwood-Nam-D’Agostino (GND) goodness-of-fit test (14). Internal validation was performed using 1,000 bootstrapped resamples and also generated the calibration plot. To further assess the calibration, we plotted Kaplan-Meier curves over the risk-stratified groups by the predicted probability by the nomogram. Finally, we tested the performance of the nomogram in the external validation cohort. The formula derived in the development cohort was applied to all patients in the validation cohort. To further validate the nomogram, patients were categorized into three risk groups according to the nomogram-based calculated 1-year disease-specific survival probability, and survival was compared between the three groups using log-rank test. Continuous variables were described as X±S, and categorical variables were described as a frequency (%). Student’s t-test, Chi-square test, or Fisher’s exact test were performed as appropriate. Statistical analyses were performed using SAS software (Version 9.4, SAS Institute, Cary, NC, USA), and P<0.05 was considered statistically significant.

**Results**

**Patient characteristics**

During the study period, 113 patients underwent pancreatectomy following neoadjuvant chemoradiation therapy. Of these, 64 were men and 49 were women. The mean age was 51.5±9.4 years old. A total of 42 (37.2%) patients were asymptomatic. In terms of resectability, 46 (40.7%) patients had resectable pancreatic cancer, 46 (40.7%) had borderline resectable pancreatic cancer, and 21 (18.6%) had locally advanced pancreatic cancer. The most common location of the tumor was the pancreatic head (86/113, 76.1%), followed by the pancreatic body/tail (27/113, 23.9%). Pylorus-preserving pancreaticoduodenectomy (PPPD) was performed in 83 (73.5%) patients, distal pancreateopancreatectomy (DPS) in 27 (23.9%) patients; and total pancreatectomy (TP) in 3 (2.7%) patients. Combined venous vascular resection including superior mesenteric vein (SMV)-splenic vein (SV)-portal vein (PV) confluence was performed in 40 (35.4%) patients. The maximum diameter of the resected pancreatic cancer was 2.1±1.5 cm. Lymphovascular and perineural invasion were observed in 10 (8.8%) and 50 (44.2%) patients, respectively. Lymph node metastasis was found in 36 (31.9%) patients. The median disease-free survival and disease-specific survival was 11.0 [95% confidence interval (95% CI): 8.8–13.2] months and 29.0 (95% CI: 1.6–38.4) months, respectively. Twenty-seven (23.9%) patients were found to have cancer-related mortality within 1 year postoperatively.

**Preoperative factors to predict 1-year disease-specific survival**

In univariate analysis, age (P=0.0165), serum CA 19-9 (after neo-Tx) following neoadjuvant treatment (P=0.0023), and preoperative image-based tumor size (P=0.0469) were statistically significant preoperative parameters to predict less than 1-year disease-specific survival in patients with pancreatic cancer who underwent pancreatectomy following neoadjuvant treatment. Subsequent multivariate analysis showed that age [hazard ratio (HR)=1.062, 95%
CI: 1.008–1.118, P=0.024] and CA 19-9 ≥34 U/mL (after neo-Tx) following neoadjuvant treatment (HR=2.863, 95% CI: 1.123–7.300, P=0.0276) were independent prognostic factors to predict less than 1-year disease-specific survival (Table 1).

**Developing nomogram to predict 1-year disease-specific survival and model performance**

Of the 113 patients, the variables of “preoperative serum CA 19-9 after neoadjuvant treatment” were missing in 9 patients, we excluded this data and 104 patients were finally analyzed. Based on the complete survival-specific data and preoperatively available parameters of the 104 patients, the nomogram to predict 1-year disease-specific survival of the patients who underwent pancreatocutomy following neoadjuvant treatment for pancreatic cancer was created (Figure 1). The nomogram was created based on 5 clinically detectable preoperative parameters: age (year), symptom (no/yes) on initial diagnosis, image-based tumor size at initial diagnostic stage (cm), preoperative serum CA 19-9 (after neo-Tx) (<34/≥34 U/mL), and planned surgery [PD (PPPD)/DPS/TP]. The model performance was quantified via Harrell’s c-statistics, which was 0.736 [standard error (SE) =0.064].

**Table 1** Univariate and multivariate analyses showing association between preoperative clinical parameters and 1-year disease-specific survival

| Variables                  | Univariate analysis | Multivariate analysis |
|----------------------------|---------------------|-----------------------|
|                            | HR                  | 95% CI                | P         | HR                  | 95% CI                | P         |
| Age (year)                 | 1.058               | 1.101–1.108           | 0.0165    | 1.062               | 1.008–1.118           | 0.0241    |
| Sex, male                  | 0.858               | 0.398–1.848           | 0.6947    | 2.019               | 0.692–5.891           | 0.1986    |
| BMI (kg/m²)                | 0.861               | 0.740–1.002           | 0.0538    | 1.243               | 0.800–1.931           | 0.3336    |
| Symptom, yes               | 1.954               | 0.826–4.625           | 0.1273    | 2.302               | 0.966–5.489           | 0.0612    |
| Weight loss, yes           | 1.206               | 0.487–2.989           | 0.6855    | 1.921               | 0.903–4.087           | 0.0903    |
| Jaundice, yes              | 0.598               | 0.268–1.331           | 0.2076    | 1.414               | 0.571–3.504           | 0.4541    |
| Abdominal pain, yes        | 1.206               | 0.487–2.989           | 0.6855    | 2.019               | 0.692–5.891           | 0.1986    |
| Dyspepsia, yes             | 0.987               | 0.406–2.398           | 0.9761    | 0.703               | 0.182–2.721           | 0.6103    |
| ASA score                  | 3 vs. 1             | 0.703                 | 0.182–2.721 | 0.6103    | 2.019               | 0.692–5.891           | 0.1986    |
|                           | 4 vs. 1             | NA                   | NA       | 1.243               | 0.800–1.931           | 0.3336    |
|                           | NA                  | NA                   | NA       | 1.679               | 0.754–3.739           | 0.2044    |
| CA19-9 (initial diagnosis) | 2.302               | 0.966–5.489           | 0.0612    | 2.640               | 1.001                 | 1.000–1.001           | 0.0023    |
| CA19-9 (after neo-Tx)      | 1.450               | 1.005–2.093           | 0.0469    | 1.243               | 0.800–1.931           | 0.3336    |
| CA19-9 (after neo-Tx) (≥34) (U/mL) | 1.450               | 1.005–2.093           | 0.0469    | 1.243               | 0.800–1.931           | 0.3336    |
| Tumor size on imaging (cm) | 1.450               | 1.005–2.093           | 0.0469    | 1.243               | 0.800–1.931           | 0.3336    |
| Smoking, yes               | 1.060               | 0.485–2.314           | 0.8841    | 1.060               | 0.485–2.314           | 0.8841    |
| Tumor location             | 1.679               | 0.754–3.739           | 0.2044    | 1.679               | 0.754–3.739           | 0.2044    |
| Initial resectability      | 1.679               | 0.754–3.739           | 0.2044    | 1.679               | 0.754–3.739           | 0.2044    |
| Predicted surgery          | PD/PPPD             | reference            |           | PD/PPPD             | reference            |           |
|                           | DPS                 | 1.751                 | 0.780–3.928 | 0.1745    | 1.922               | 0.697–5.285           | 0.2068    |
|                           | TP                  | 3.726                 | 0.493–28.134 | 0.2022    | 1.522               | 0.179–12.904          | 0.7003    |

BMI, body mass index; ASA, American Society of Anesthesiologist; NA, not available; CA, carbohydrate antigen; neo-Tx, neoadjuvant treatment; Resectability 1, resectable; Resectability 2, borderline resectable; Resectability 3, locally advanced; PD, pancreaticoduodenectomy; PPPD, pylorus-preserving PD; TP, total pancreatectomy; HR, hazard ratio; 95% CI, 95% confidence interval.
Validation of nomogram

Model performance for discrimination and calibration was assessed. Discrimination was validated using the bootstrap validation with 1,000 resamples. The calibration plot (Figure 2) showed good agreement between actual and predicted survival probabilities, and the GND goodness-of-fit test showed that the model was well calibrated ($\chi^2=8.24$, $P=0.5099$). To further assess the discriminating capability of the nomogram, according to nomogram-based estimated 1-year disease-specific survival probability, patients were categorized into three risk groups as follows: the low-risk group comprised patients with a nomogram-based calculated 1-year disease-specific survival probability of <33.3%; the intermediate-risk group, 33.3%–66.6%; and the high-risk group, >66.6%. Statistically significant differences in the survival rates were found among the three groups ($P<0.001$, Figure 3).

External validation to assess nomogram performance in predicting 1-year disease-specific survival

Institutional data of the patients who underwent resection following neoadjuvant treatment were obtained from five international hospitals. A total of 84 patients were used for
external validation to assess the nomogram performance. Patients’ age (61.4±9.6 vs. 66.9±9.7, P=0.0001) and serum CA 19-9 (after neo-Tx) (≥34 U/mL) following neoadjuvant treatment [48 (46.2%) vs. 64 (74.4%), P<0.0001] were statistically significant between the study cohort and the validation cohort (Table 2). Moreover, when correlating actual disease-specific survival and calculated 1-year disease-specific survival, there was statistically significant differences according to the calculated probability of 1-year survival among the three groups (P=0.044, Figure 4).

**Discussion**

Pancreatic cancer is the fourth leading cause of cancer-associated mortality, with approximately 43,090 deaths recorded in 2017 in the United States (15). The median survival is 26 months, and the 5-year survival rate is 30% (6). Margin-negative pancreatectomy followed by adjuvant chemotherapy is the standard treatment modality for resectable pancreatic cancer. However, only 15%–20% of the patients are eligible for resection at the time of diagnosis. In addition, patients who underwent margin-positive pancreatectomy show poor survival outcomes similar to those with locally advanced and unresectable pancreatic cancer (6). Therefore, neoadjuvant treatment has been utilized for borderline resectable and locally advanced (unresectable) pancreatic cancer to improve eligibility for surgical resection, with the ultimate aim to optimize the treatment outcomes and minimize the risk of treatment-related morbidity.

Studies (16,17) have shown the importance of neo-adjuvant treatment in pancreatic cancer. A recent meta-analysis (18) of studies reporting median overall survival by intention to treat in patients with resectable or borderline resectable pancreatic cancer treated with or without neoadjuvant treatment showed that neoadjuvant treatment appears to improve overall survival (18.8 months vs. 14.8 months) despite lower overall resection rates (66.0% vs. 74.0%).

![Figure 3](image.png)

**Figure 3** Correlation of actual patients’ survival according to calculated probability of 1-year disease-specific survival (internal validation). Low-risk group comprised patients with nomogram-based calculated 1-year disease-specific survival probability of <33.3%; intermediate-risk group, 33.3%–66.6%; and high-risk group, >66.6%. The survival rates were significantly different among the three groups (P<0.001).

| Variables | Study cohort (N=104) | Validation cohort (N=86) | P |
|-----------|---------------------|------------------------|---|
| Age (year) | 61.4±9.6 | 66.9±9.7 | 0.0001 |
| Symptom, yes | 66 (63.5) | 58 (67.4) | 0.5663 |
| Preoperative serum CA19-9 (after neo-Tx) (≥34) (U/mL) | 48 (46.2) | 64 (74.4) | <0.0001 |
| Tumor size on imaging (cm) | 2.9±0.9 | 2.9±1.3 | 0.6682 |
| Planned surgery | | | 0.6266 |
| PD/PPPD | 79 (76.0) | 60 (69.8) | |
| DPS | 23 (22.1) | 24 (27.9) | |
| TP | 2 (1.9) | 2 (2.3) | |
| Death (event) | 22 (21.2) | 12 (14.0) | 0.1975 |

CA, carbohydrate antigen; neo-Tx, neoadjuvant treatment; PD, pancreaticoduodenectomy; PPPD, pylorus-preserving PD; DPS, distal pancreatectosplenectomy; TP, total pancreatectomy; Categorical and continuous variables were analyzed via Chi-square test or Fisher’s exact test and student’s t-test, respectively.
In addition, Choi et al. (19) conducted a prospective randomized controlled trial on the oncologic benefits of neoadjuvant treatment in borderline resectable pancreatic cancer and showed that neoadjuvant treatment provided oncologic benefit over upfront surgery in patients with borderline resectable pancreatic cancer. Even in the metastatic setting, FOLFIRINOX and nab-paclitaxel-gemcitabine (20,21) have shown a survival advantage over previously standard gemcitabine monotherapy; thus, they have become the standard treatment option in patients with good performance status. Based on these clinical evidences, neoadjuvant treatment is expected to be actively used in managing pancreatic cancer.

Deciding on the necessity of radical pancreatectomy following neoadjuvant for pancreatic cancer can be complicated. Most patients who underwent neoadjuvant treatment have borderline resectable and locally advanced pancreatic cancer and usually require combined vascular resection with extensive surgical dissection for margin-negative resection. Several studies have reported that extended radical pancreatectomy is associated with increased morbidity and mortality (10,12). In addition, except for resectable pancreatic cancer, other tumors are biologically aggressive. Thus, patients are at high risk of early recurrence and early cancer-related mortality despite successful curative surgical resection. Therefore, the patients and their families should be well-informed of these risks before surgery. Moreover, the balance between potential risk and oncologic benefit should be considered in deciding the need for surgery. However, many researches on survival analysis in patients with pancreatic cancer who underwent surgery following neoadjuvant treatment are based on pathological examinations after surgical intervention, such as tumor grade (22), pathological complete response (23), lymph node status (24), perineural invasion (25), and margin-status (6). The response of neoadjuvant chemotherapy is thought to be an important prognostic factor in pancreatic cancer with neoadjuvant treatment (23). However, in our clinical practice, it is true that response to neoadjuvant treatment cannot be always measured objectively due to inflammatory changes and previous biliary stent. Therefore, in order to develop nomogram to be applied in every case of pancreatectomy following neoadjuvant treatment, we did not include the response of neoadjuvant chemotherapy, but include preoperative serum CA 19-9 after neoadjuvant treatment. Methods to preoperatively predict the survival outcome for these patients are thus important in clinical practice.

In the present study, only preoperative clinical parameters, such as, age, symptom, image-based tumor size at diagnosis, preoperative serum CA 19-9 after neoadjuvant treatment, and planned surgery, were considered in developing a nomogram (internet calculator; http://40.121.207.11:8080/home1.jsp) to predict 1-year survival outcomes after pancreatectomy following neoadjuvant treatment for pancreatic cancer. In univariate analysis, variables with P<0.2 were age, body mass index (BMI), symptom, abdominal pain, CA19-9 (after neo-Tx), CA19-9 (≥34 U/mL) tumor size, initial resectability, and predicted surgery. The three parameters (symptom, tumor size, and planned surgery) were of no significance in multivariate analysis but were used in the development of the nomogram. The reason for this was to consider the concept of events per predictor variable (EPV) when constructing the nomogram and to include variables with clinical importance among variables with P<0.2. Although there are some variables that are not significant when multiple logistic regression analysis is performed, the reason why we built the prediction model with these five variables was because we wanted to find a model with better overall model performance than the significance of each variable. We selected this model as the final model because it
showed the best performance when modeled using these five variables in this data. Harrell’s c-index was 0.684 (SE=0.064) when the nomogram was created with age, tumor size, and CA19-9 after neoadjuvant treatment with  \( P<0.05 \) in univariate analysis and when the nomogram was created with age, abdominal pain, and CA19-9 after neoadjuvant treatment, which were selected by the stepwise selection method, the Harrell’s c-index was 0.711 (SE=0.064). Internal validation of the current nomogram via Harrells’ c-statistics (0.736) and the calibration plot demonstrated its quite acceptable clinical accuracy. In addition, external validation of the clinical availability and feasibility of the current nomogram in predicting survival outcome in patients who underwent pancreatectomy following neoadjuvant treatment was conducted using dataset from 5 international institutions. Similar to the primary observation from the development dataset, the predicted 1-year disease-specific survival was successfully correlated with actual survival outcomes in external validation set, which is among the strong points of the current nomogram. However, when correlating actual disease-specific survival and calculated 1-year disease-specific survival, the survival curves between high-risk group and intermediate-risk group crossed in the external validation data (Figure 4). We think the statistical power is a bit lower due to the small number of event cases in the high-risk group in the external data. Even though the \( P \)-value of external validation set increased compared to development set, considering \( P \)-value less than 0.05, the prediction effect seems to be useful. It was thought to be meaningful that we performed external validation with multinational data to evaluate generalizability as well as internal validation for our prediction model.

However, when we tried the GND goodness-of-fit test in the survival setting, the number of events was too small to calculate, therefore, the \( P \)-value calculation for accuracy between predicted value and actual value seems to be a limitation in this study. The accuracy and clinical application of this nomogram should be re-assessed in prospective large volume-patient cohorts. Particularly, in the medical field influenced by the 4th Industrial Revolution, such methods are predicted to be developed via data-based artificial intelligent algorithms. The continued increase in available data and emerging computational methods indicate that precision public health is highly likely to be available in the future (30).

**Conclusions**

The nomogram to predict 1-year disease-specific survival for patients with pancreatic cancer who are eligible for surgery following neoadjuvant treatment showed accuracy and clinical feasibility for decision-making in the management of pancreatic cancer.

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**Footnote**

*Conflicts of Interest:* The authors have no conflicts of interest to declare.

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