Volatility Modelling for Tourism Sector Stocks in Borsa Istanbul

Gulsah Gencer Celik*

Meslek Yüksekokulu, Çocuk Gelişimi Programı, Beykent Üniversitesi, İstanbul, Turkey. *Email: gulsahg@beykent.edu.tr

Received: 15 March 2020 Accepted: 21 April 2020 DOI: https://doi.org/10.32479/ijefi.9811

ABSTRACT

This paper examines the volatility of the tourism sector in Borsa İstanbul in Turkey, paying special attention to the role of exchange rate exposure in the process. The GARCH, BJR (TARCH) and EGARCH models are employed to estimate the volatility in the stock returns of Turkish tourism firms using daily data from 2 January 2002 to 13 April 2020. The results suggest that: (i) Compared to the GARCH and GJR model results, the EGARCH model provides valuable information on the volatility of returns in tourism sector and on the impact of exchange rate on stock returns; (ii) the impact of exchange rate risk on stock returns is significant and positive for 3 tourism firms and negative for 2 firms; (iii) the findings on volatility of stock returns indicate that the time-dependent components of volatility is clearly more important than the time-independent component of volatility in predicting current volatility; (iv) the volatility of stock returns are highly persistent and the volatility at time t is more sensitive to past period volatility than past surprises in the market; (v) surprisingly, while there is no leverage effect, shocks have asymmetric effect on volatility implying that the impact of negative news do not outweigh positive news (or the impact of positive news on volatility is higher than the impact of negative news in the market).

Keywords: Turkish Tourism Industry, Volatility, Foreign Exchange Rate Risk, Stock Returns, ARMA, GARCH, GJR(TARCH), EGARCH Model

JEL Classifications: G1, N2, C5

1. INTRODUCTION

Volatility is used to measure the dispersion of returns in the stock prices. The volatility of stock returns is affected by a large number of risk factors such as political instability, economic fundamentals, government budget deficits, economic policy changes, firm-specific factors, and so on. For the tourism industry, exchange rate exposure can be considered as the most important risk factor that affects the stock return of tourism firms. For this reason, this study aims to investigate the volatility of the tourism firms’ stock returns listed in Borsa İstanbul, paying special attention to the role of exchange rate exposure in the process. The subject matter is important for several reasons and of great interest to researchers on the subject, tourism firm managers, policymakers, and portfolio managers. For policymakers, being an important foreign exchange generating sector of the Turkish economy, the tourism sector plays a vital role in the balance of payments of the economy. For managers, the subject is important because the tourism sector is very sensitive to external shocks and especially the exchange rate shocks.

The impact of exchange rate exposure on stock returns might have a positive or negative effect. In this sense, we can identify four main channels through which exchange rate risk affects stock returns (Kasman et al., 2011; Olugbode et al., 2014): (1) According to the intertemporal capital asset pricing model and the Arbitrage Pricing Theory (APT), investors require additional compensation for bearing the risk of exchange rate changes and hence exchange rate sensitivities exert a significant impact on the common stocks of tourism firms; (2) Exchange rate exposure plays a vital role in the profitability of firms by influencing the value of a firm. Fluctuations in exchange rates can affect the value of the firm (Vardar et al., 2008; Kasman et al., 2011; Fauziah et al., 2015; Dornbusch and Fischer, 1980), through influencing the cash flows of multinational firms, importers, exporters, and also purely domestic firms (Hyde, 2007; Lin, 2012); (3) Maturity mismatch between the assets and liabilities of tourism firms and unexpected change in interest and exchange rates are considered as the key factors that lead to increase the risk exposure of the tourism firms; (4) the revenues, costs, and profitability of tourism firms are directly influenced.
by the unexpected changes in exchange rates (Saunders and Yourougou, 1990). Depending on the net foreign positions of a firm’s balance sheet, the unexpected movements in exchange rates can lead to gains or losses. For example, when foreign currency-denominated liabilities exceed foreign currency denominated assets, the depreciation of the local currency may lead to damage in the firms’ balance sheet, and the deterioration of firms’ equity may result in a decline in the tourism firms’ stock return.

In recent years, it is often argued that volatility, especially in financial markets, has increased in line with financial globalization. Financial globalization intensifies volatility during the periods of high uncertainty, increases instability in a country facing external shocks and become a destabilizing factor in the economy (Çelik, 2019; Stiglitz, 2004; Cordella and Rojas, 2017; Kose et al., 2009; Umutlu et al., 2010). In particular, the higher volatility of Turkish Lira during crisis periods of 2008-2009 and 2013 has exerted an important effect on stock returns through increasing uncertainty, affecting exports, imports, foreign direct investment, and portfolio investment decisions.

In light of these discussions, this study models volatility of stock returns series of tourism firms in Turkey using ARMA-GARCH type models. The GARCH, GJR (or TARCH) and EGARCH models will be estimated using daily data of six tourism firms listed in Borsa İstanbul over the period of 2003-2020. The rest of this study is organized as follows. Section 2 reviews the relevant empirical literature on the relationship between stock returns of tourism firms and the exchange rate. Section 3 provides the data subject to empirical analysis and introduces the empirical model employed in this study. Section 4 presents the findings obtained from estimating the GARCH, GJR (or TARCH) and EGARCH models. Section 5 concludes.

2. LITERATURE REVIEW

Over time, a vast amount of literature is accumulated on estimating of volatility in stock returns (Song, 1994; Mansur and Elyasiani, 1995; Flannery et al., 1997; Engle et al., 1990; Elyasiani and Mansur, 1998; Sehgal and Agrawal, 2017; Yamak et al., 2018; Çelik, 2019; Kasman et al., 2011; Olugbode et al., 2014, etc.). However, the number of empirical studies that investigate volatility in the tourism sector is limited, and most of these studies are undertaken at a sectoral level (Gokmenoglu and Hadood, 2019; Hsiao, 2017; Chang et al., 2013; Lee and Jang, 2010; Mohapatra, 2017).

A review of the empirical literature indicates that there are only a few studies on the relationship between exchange rate exposure and stock return volatility in the tourism industry, particularly at a firm level. Chang et al. (2013) examine the size effects of volatility spillovers for firm performance and exchange rates in the Taiwan tourism industry using BEKK-AGARCH and VARMA-AGARCH models. The authors find that there are size effects on volatility spillovers from the exchange rate to firm performance, and there is a negative correlation between exchange rate returns and stock returns.

In another study, Gokmenoglu and Hadood (2019) analyzed the volatility spillover between foreign exchange rate and tourism firm stock returns in China utilizing the BEKK-GARCH model. The results of the study indicate that there is bidirectional long-term spillover volatility between the variables under investigation. In their study for the US tourism firms, Obi et al. (2015) concluded that U.S. tourism firm stock performance had only an adverse long-run association with expected risk proxied by S&P 500 implied volatility. Using the ARIMA model, Hsiao (2017) examines the effect of fourteen foreign currencies on twelve selected Taiwanese tourism firms’ profitability. Results showed that return on assets (ROA) and return on equity (ROE) are differently and significantly affected by the exchange rate of Taiwanese currency against foreign currency fluctuations. By utilizing BEKK-AGARCH and GJR-AGARCH models, Chang et al. (2013) found out that there is a bidirectional size effect of long-term volatility spillovers between the Taiwanese foreign exchange rate (against U.S. dollar and Chines Yuan) and large tourism firm stocks, while long-run volatility of small firms’ stocks was only affected by Japanese long-term exchange rate volatility.

Review of the empirical studies related to Turkish tourism firms indicates that the empirical studies on the subject mainly aimed at measuring the financial performance of Turkish tourism companies with the help of financial ratios (Özçelik and Kandemir, 2015; Karadeniz and Iskenderoğlu, 2011; Erdoğlu, 2014; Güdük, 2018; Ergül, 2014).

Using a different methodology, Doğukanlı et al. (2010) investigated the exchange rate sensitivity of the main and sub-sector stock indices in the Borsa İstanbul in terms of Dollar and Euro currencies. They have used Johansen cointegration analysis. The results showed that there is a cointegration relationship between sectoral stock indices and exchange rates. Using a similar methodology, Soyaslan (2019) examined the relationship between the exchange rate and the BIST tourism stock index using cointegration analysis. She found out that there is a long-run cointegration relationship between the exchange rate and the BIST tourism index.

In their study, Kutlu and Karakaya (2019) attempted to investigate the volatility of the Borsa İstanbul Tourism Index with the two-stage Markov Regime Change Autoregressive Conditionally Changing Variance model. The study was conducted between the periods of 02 May 2003 and 14 September 2018 in three periods, before the 2008 financial crisis, the 2008 crisis, and after the 2008 financial crisis. According to the results obtained with Markov Regime Change Autoregressive Conditional Variable Variance Model, Tourism index volatility could not return to the pre-crisis period. With the effect of the global crisis, the tourism index has volatility in all three periods, and volatility in the post-crisis period is higher than in the pre-crisis period.

Different from previous studies, Kandil et al. (2020) examined the effect of the exchange rate and interest rate on equity profitability for six tourism companies listed in the BIST. In their analysis, the long-term effect of the exchange rate and the interest rate on equity profitability of firms was examined by the Maki Cointegration test,
and the direction and coefficient of the effect were determined by the DOLS estimator. The findings show that both systematic risk factors have a negative effect on the profitability of the listed tourism companies.

The review of the empirical literature on the volatility of tourism stock returns shows that almost all studies on the subject are carried out at an aggregate level, and they differ significantly in terms of methodology they employed. Furthermore, the empirical studies that aimed at estimating volatility using firm-level data in the tourism sector are very limited in number.

**3. DATA AND METHODOLOGY**

**3.1. Data**

This study aims to examine the volatility spillover between foreign exchange rates and tourism stock returns in Turkey. The study employed the data obtained from the Finnet Data Delivery System and electronic data delivery system of Turkish Central Bank of Turkey. Stock prices for a sample of five Turkish tourism firms’ stocks listed on the Borsa Istanbul (BIST) are collected and calculated from the Finnet Data Delivery System. It is daily data for the period 02 January 2002-2013 April 2020 with 4593 observations. The tourism firms that their data analysed includes AYCES, MAALT, MARTI, METUR, PKENT and TEKTU. These firms are chosen due to data availability. Dolar Exchange rate of Turkish Lira is obtained from the electronic data delivery system of Turkish Central Bank of Turkey of the sample period. Returns on exchange rates (ER) and tourism firms stock prices (R) are calculated by taking the first difference in log prices as

\[ R_t = \left[ \ln(P_t) - \ln(P_{t-1}) \right] \times 100 \]

where \( P_t \) and \( P_{t-1} \) are daily closing prices at time \( t \) and \( t-1 \) respectively.

Tablo 1 presents the descriptive statistics for each of the five tourism firms’ stocks and Dolar/Turkish Lira exchange rate returns (ER). While the mean returns for tourism firms’ stock returns are positive, ranging from a 0.0092 (METUR) to 0.0708 (MAALT), volatilities in stock returns for tourism firms range from 3.0314% to 3.7393%. However, the volatility of the Dolar exchange rate of domestic currency is relatively small, with a standard deviation of 0.895% compared to volatilities in stock returns. Furthermore, the results in Table 1 shows that all series subject to empirical analysis has a highly skewed (skewed to the left) and leptokurtic distribution rather than the normal distribution. The null hypothesis of normality is rejected for all series by the Jarque–Bera normality test.

**3.2. Unit Root Test**

For time-series data, it is important to test for the stationarity of the data since non-stationary regressors may invalidate most of the standard empirical results (Engle and Granger 1987; Enders 2015). In this study, we used the Augmented Dickey-Fuller (ADF) test and Phillips-Perron tests to determine the level of integration of the variables of interest (see Dickey and Fuller, 1981; Phillips and Perron, 1988). For each of the variables subject to empirical analysis, both ADF and PP statistics were calculated for the series including, no intercept and no trend, intercept, and intercept and trend in the underlying Phillip-Perron and Dickey-Fuller regressions. Table 2 presents the unit root test results. Inspection of the results Table 2 shows that the hypothesis of a unit root for all series is rejected at 1% level of significance, indicating that all series are stationary.

**Table 1: Descriptive statistics, 02 January 2002-2013 April 2020**

|         | AYCES | MAALT | MARTI | METUR | PKENT | TEKTU | ER  |
|---------|-------|-------|-------|-------|-------|-------|-----|
| Mean    | 0.0627| 0.0708| 0.0172| 0.0092| 0.0695| 0.0281| 0.033|
| Median  | 0.0000| 0.0000| 0.0000| 0.0000| 0.0000| 0.0000| -0.017|
| Maximum | 19.6115| 18.2322| 20.2524| 20.1637| 19.2372| 19.8851| 14.706|
| Minimum | -17.6456| -19.5567| -20.0671| -21.8002| -21.3093| -21.7065| -11.931|
| Std. dev.| 3.0314| 3.2035| 3.1348| 3.4206| 3.7393| 3.5740| 0.895|
| Kurtosis | 0.5731| 0.6756| 0.3962| 0.5570| 0.8467| 0.2815| 1.123|
| Skewness | 9.1135| 9.6691| 8.3008| 10.2680| 8.5221| 8.2628| 32.692|
| Jarque-Bera | 7402.56| 8859.22| 5496.36| 10028.94| 6383.10| 5359.92| 169651.70|
| Probability | 0.0000| 0.0000| 0.0000| 0.0000| 0.0000| 0.0000| 0.0000|
| Sum     | 288   | 325   | 79    | 41    | 319   | 129   | 153  |
| Sum sq. dev. | 42189| 47115| 45115| 52078| 64193| 58644| 3675|
| N. of observations | 4592| 4592| 4592| 4592| 4592| 4592| 4592|

**Table 2: Unit root test results**

|         | ADF (τ) | ADF (τ_1) | ADF (τ_2) | PP (τ) | PP (τ_1) | PP (τ_2) | PP (τ_3) |
|---------|---------|-----------|-----------|--------|----------|----------|----------|
| AYCES   | -43.4855 | -43.4829  | -43.4665  | -65.0387| -65.0332 | -65.0396 | -65.0396 |
| MAALT   | -64.1807 | -64.1752  | -64.1589* | -65.0387| -64.4961 | -64.4944 | -64.4944 |
| MARTI   | -68.0646 | -68.0577* | -68.0707* | -65.0387| -64.0909 | -68.0641 | -68.0679 |
| METUR   | -61.8159 | -61.8119* | -61.8226* | -65.0387| -62.1397 | -62.1348 | -62.1461 |
| PKENT   | -51.6484 | -51.6447* | -51.6251* | -65.0387| -70.6785 | -70.6745 | -70.6315 |
| TEKTU   | -66.0488 | -66.0437* | -66.0526* | -65.0387| -66.2628 | -66.2628 | -66.3024 |
| ER      | -62.6675 | -62.7364* | -62.5932* | -65.0387| -62.5454 | -62.5106 | -62.5065 |

ADF and PP refer to the Augmented Dickey and Fuller (1981) and Phillips and Perron (1988) unit root tests. The lag lengths in the ADF and PP regressions are determined by the Schwarz Information Criteria (SIC). Asterisks (*,**,***) shows the 1%, 5%, and 10% the level of significance
3.3. Empirical Model
To examine the volatility of the tourism sector in Borsa İstanbul in Turkey, paying special attention to the role of exchange rate exposure in the process, the GARCH-type (Generalized Autoregressive Conditional Heteroscedasticity) models are estimated. The reason for choosing the GARCH-type models is that volatility and changing variance are characterized as the nature of high frequency economic and financial time series. The GARCH-type models provide a relevant framework to model the presence of heteroscedasticity as a conditional variance. Since the GARCH models treat conditional heteroscedasticity as a variance to be modeled rather than as a problem to be corrected, the GARCH-type models can be used to estimate the relationship between exchange rate and stock returns of tourism firms as a conditional variance process. In empirical studies, the GARCH model (Bollerslev, 1986), the BJG (or TARCH) model (Glosten et al., 1993), and the EGARCH model (Nelson, 1991) are widely used models in modeling the volatility in return series. Three GARCH models will be estimated in this paper, namely the GARCH, GJR (or TARCH), and EGARCH. The model specifications of these models can be briefly explained as follows. These models have two components, conditional mean and conditional variance specifications.

3.3.1. Conditional mean specification
\[ R_t = a_0 + a_1 R_{t-1} + a_2 R_{t-2} + 0 ER_t + u_t - \delta u_{t-1} \]
(1)
where \( R_t \) represents daily stock price returns and \( ER_t \) is percentage change in daily exchange rate at time \( t \). \( u_t \) represents normally distributed error terms with mean zero and the conditional variance of \( h_t \).

3.3.2. Conditional variance specification
The GARCH model (Bollerslev, 1986) treats the conditional variance as a function of its own lags as well as lagged shocks to stock price returns. The conditional variance equation for GARCH(1,1) model can be stated as (Enders, 2015):
\[ h_t = \omega + \alpha u_{t-1}^2 + \beta h_{t-1} \]
(2)
where \( h_t \) is the conditional variance, namely a one-period ahead estimate (or forecast) of the conditional variance based on past information, \( \omega \) is a constant term, and \( u_{t-1}^2 \) measures the shocks in volatility. \( h_{t-1} \) is the forecasted variance from yesterday. To ensure that the conditional variance is positive, the parameter estimates should be as \( \omega > 0, \alpha \geq 0, \beta \geq 0 \). Furthermore, the necessary and sufficient condition that \( \alpha + \beta < 1 \) should hold for the existence of the second moment of \( u_t \) for GARCH(1,1).

The GARCH model is a widely used model in practice since it is possible to model very complex conditional variance processes using only fewer parameters. However, it does not incorporate the asymmetric volatility. In other words, the GARCH model enforces the asymmetric response of volatility to positive and negative shocks of equal magnitude.

3.4. GJR-GARCH Model
The GJR or TARCH model developed by Glosten et al. (1993) incorporates asymmetric volatility. The GJR model is an extension of the GARCH model with an additional term added to account for possible asymmetries. The advantages of the GJR model are that the asymmetric effects of positive and negative shocks are directly modeled, simpler to implement in practice, and provides better performance in forecasting volatility (Liu and Hung, 2010). The conditional variance of the GJR (or TARCH) model is given by:
\[ h_t = \omega + \alpha u_{t-1}^2 + \beta h_{t-1} + \gamma u_{t-1}^2 D_{t-1} \]
(3)
\[ D_{t-1} = 1 \text{if } e_{t-1} < 0 \]
\[ D_{t-1} = 0 \text{if } e_{t-1} \geq 0 \]
where \( h_t \) is the conditional forecasted variance, \( \omega \) is the intercept for the variance, \( u_{t-1}^2 \) is the variance that depends on previous lag error terms, \( h_{t-1} \) is the forecasted variance from yesterday and \( D_{t-1} \) is a dummy variable that takes 1 for negative shocks and 0 zero for positive shocks. To make sure that the variance, \( h_t > 0 \), the sufficient conditions involve \( \omega > 0, \alpha \geq 0, \beta \geq 0, \text{ and } \alpha + \gamma \geq 0 \).

The coefficients in Equation 3 provide rich interpretation related to the volatility of returns. The \( \gamma \) parameter provides information about a possible asymmetric effect in data: If \( \gamma = 0 \), there is no asymmetric volatility. If \( \gamma > 0 \) negative shocks will increase risk (volatility) more than positive shocks of the same magnitude, and if \( \gamma < 0 \), positive shocks increase the volatility more than negative shock. The total effect of a negative shock is equal to \( (\alpha + \gamma) \) and a positive shock is equal to \( \alpha \). \( \beta \) captures the effects of persistence of shocks on volatility of returns.

3.5 EGARCH Specification of the Conditional Variance
The EGARCH model developed by Nelson (1991) provides an alternative specification for the conditional variance. This specification several important advantages (superior features) compared to GARCH and GJR conditional variance models. Because the logarithm of conditional volatility, \( \ln (h_t) \), is modeled, the conditional variance is always positive, and there is no need to artificially impose non-negativity constraints on the model parameters. The condition that \( |\beta| < 1 \) is a sufficient condition for the existence of moments, for consistency, and for asymptotic normality of the EGARCH(1,1) estimators (Chang et al., 2014). The EGARCH variance equation can be written as:
\[ \ln (h_t) = \omega + \alpha \left( \frac{u_{t-1}^2}{\sqrt{h_{t-1}}} \right) + \beta \ln (h_{t-1}) \]
(4)
where \( \ln (h_t) \) represents the logarithm of conditional variance at time \( t \), \( \omega \) is the intercept for the variance, \( \beta \) is the coefficient for the logged GARCH term \( \ln (h_{t-1}) \) indicating the persistence of shocks, \( \gamma \) is the scale of the asymmetric volatility. If \( \gamma < 0 \) and significant, it indicates the presence of leverage effect. The leverage effect refers to the negative correlation between an asset return and its volatility. In this sense, \( \gamma < 0 \) shows that bad news
(or negative shocks) generates larger volatility than positive shocks. However, if $\gamma$ is not significant, then there is no asymmetric volatility. If $\gamma > 0$ and significant, this means that positive shocks increase the volatility more than negative shocks. The coefficient $\alpha > 0$ represents the tendency of shocks to persist and shows the presence of volatility clustering. That is, volatility tends to rise when $\frac{u_{t-1}}{\sqrt{h_{t-1}}}$ is larger and vice versa.

### 4. RESULTS

This section provides the empirical findings on the impact of the exchange rate on Turkish tourism firms’ stock returns obtained from estimating the GAR, TARCH, and EGARCH models given in Equations 2, 3, and 4. As mentioned in section 3, the presence of heteroscedasticity or clustering of observations in error terms invalidates the standard econometric results. In such cases, the GARCH-type models provide a very useful framework to model volatility in the series. Before estimating the GARCH-type models, we first formulated and estimated the suitable ARMA model given in Equation 1 for the stock returns of tourism firms and then tested for the presence of the clustering of error terms. Table 3 presents the findings on ARMA models. The type of ARMA models is determined by Schwarz Information Criteria. The results given in Table 3 show that the exchange rate has a positive and significant effect on tourism firms’ returns for only two firms out of six cases. The diagnostic statistics related to ARMA models indicate that

The estimation results related to the conditional variance equation are given in Table 4. The results indicate that non-negativity conditions, the covariance stationarity, and stability conditions are satisfied since the $\omega$, $\alpha$, and $\beta$ coefficients are positive, and $\alpha + \beta < 1$. The results also show that the time-independent component of volatility ($\omega$) is positive and statistically significant

#### Table 3: The impact of exchange rate on tourism firms’ stock returns: The estimation results of the ARMA model

| Coefficient | AYCES | MAALT | MARTI | METUR | PKENT | TEKTU |
|-------------|-------|-------|-------|-------|-------|-------|
| $a_1$       | 0.057 (0.0509) | 0.0678 (0.0501) | 0.0189 (0.0463) | 0.0052 (0.0564) | 0.0726 (0.0456) | 0.0277 (0.0528) |
| $\theta$    | 0.0988** (0.0506) | 0.0576 (0.0540) | -0.0530 (0.0517) | 0.1035*** (0.0587) | -0.0955 (0.0621) | 0.0117 (0.0590) |
| $a_1$       | 0.0522* (0.0149) | 0.0580* (0.0150) | -0.0455 (0.1930) | -0.2715** (0.1293) | 0.6791* (0.0984) | 0.0277 (0.0528) |
| $\alpha$    | 0.0722* (0.0148) | 0.0580* (0.0150) | -0.0455 (0.1930) | -0.2715** (0.1293) | 0.6791* (0.0984) | 0.0277 (0.0528) |
| $\delta$    | 0.3320* (0.1221) | 0.3320* (0.1221) | 0.6791* (0.0984) | 0.0277 (0.0528) | 0.0277 (0.0528) | 0.0277 (0.0528) |

F-statistic: 12.8584* 6604.04* 8160.36* 5914.91* 8410.24* 7301.17* 0.0393

$2 \chi^2_{JB}$: 11.3432 8.6167 2.6456 11.0643 9.1103 8.4813

$2 \chi^2_{auto}$: 5.0546 5.1682 5.2959 5.6013 5.7862 5.8479 0.0393

ARCH(7): 450.94* 439.70* 278.53* 510.90* 513.77* 374.12* 0.0393

#### Table 4: The GARCH model estimates of tourism stock returns model

| Coefficient | AYCES | MAALT | MARTI | METUR | PKENT | TEKTU |
|-------------|-------|-------|-------|-------|-------|-------|
| $a_1$       | 0.0440 (0.0339) | 0.0228 (0.0417) | 0.0133 (0.0374) | 0.0017 (0.0402) | -0.0395 (0.035) | 0.0221 (0.0467) |
| $\theta$    | 0.1206* (0.0383) | 0.0622 (0.0471) | -0.0931** (0.0455) | 0.0341 (0.0395) | -0.0702*** (0.0412) | 0.0436 (0.0461) |
| $a_1$       | 0.0521* (0.0162) | -0.0617* (0.0164) | 0.5232* (0.0947) | -0.6247* (0.0868) | 0.0436 (0.0461) | 0.0436 (0.0461) |
| $\beta$     | 0.6947* (0.0103) | 0.8410* (0.0048) | 0.7798* (0.0117) | 0.6782* (0.0107) | 0.6013* (0.0110) | 0.8739* (0.0046) |

$\omega = 0.0440 \pm 0.0022 + 0.0622 + 0.0521 + 0.6947 - 0.0617 + 0.8410 + 0.7798 + 0.6782 + 0.6013 + 0.8739$. Numbers in parentheses indicate standard errors * * ** indicate the significance level at 1%, 5%, and 10% respectively
for all the tourism firms indicating that volatility is an important integral part of the return generating process in tourism industry. The coefficients on the time-dependent ARCH (α) and GARCH (β) components of volatility are positive significant for all cases. These findings provide important insight into the sources and the timing of volatility of tourism stock returns. First, the short-run persistence of shocks (new surprises) to returns (ARCH effect, α) parameter is smaller than the long-run persistence of a previous period’s forecast variance (GARCH effect, β). Second, the value of α and β parameters are very close to unity suggests that shocks to the tourism stock returns at time t have highly persistent effects implying that shocks to volatility dissipate slowly.

Table 5 presents the estimates obtained from the GJR (TARCH) model of tourism returns given in Equations 1 and 3. Examination of the mean equation shows that the impact of exchange rate risk on stock returns is exactly the same as the findings of the GARCH model. In this sense, the GARCH and GJR models provide similar results. In terms of volatility of stocks, however, as mentioned above, the GARCH model implicitly assumes that the effects of negative and positive shocks have symmetric and same effects on conditional variance (volatility) of stock returns. However, the empirical studies indicate that negative shocks (bad news or shocks that cause stock prices to decline) have a greater effect on conditional volatility than positive innovations (good news or shocks that lead stock prices to rise) of the same magnitude. In this sense, the GJR model provides further information about the importance of asymmetric effects of shocks in addition to the ARCH and GARCH effects of the GARCH model. Act of negative and positive shocks to stock return. The examination of Table 5 shows that ARCH and GARCH parameters are positive and significant for all tourism firms’ returns. This indicates that the volatility is changing by time and the impact of positive past surprises (αs) and past volatility (βs) have a highly persistent effect on current volatility (current conditional variance) of returns as in the GARCH model in Table 4. More importantly, although its magnitude is small, the parameter, γ, which represents the possible asymmetric effect in data, is negative and significant and negative for only 3 out of 6 tourism firms. Since the total effect of a negative shock is equal to (α+γ) and a positive shock is equal to α, as mentioned above, the negative, γ, parameter indicate that positive shocks increase the volatility of returns more than negative shocks of the same magnitude. β captures the effects of persistence of shocks on volatility of returns.

Table 6 present estimation results of the EGARCH model of the stock returns of tourism firms given in Equations 1 and 4. The mean equation given in Table 6 suggests that the exchange rate risk variable exerts a significant impact on stock returns of 5 tourism firms out of 6 firms. The sign of the exchange rate risk coefficients are significant and positive in 3 cases and significant and negative in 2 cases. Compared to the GARCH and GJR model results, the EGARCH model performs better in the estimation of the exchange rate stock return relationship. Evaluation of the estimates of the variance equations in Table 6 provides a number of important information about the volatility of stock returns. First, the results

Table 5: The GJR (TARCH) model of tourism stock returns model

| Coefficient | AYCES | MAALT | MARTI | METUR | PKENT | TEKTU |
|-------------|-------|-------|-------|-------|-------|-------|
| Mean equation |       |       |       |       |       |       |
| α | 0.0348 (0.0378) | 0.0318 (0.0433) | 0.0167 (0.0389) | 0.0188 (0.0454) | −0.0155 (0.0368) | 0.0259 (0.0478) |
| β | 0.1206* (0.0384) | 0.0635 (0.0472) | −0.0928** (0.0455) | 0.0358 (0.0392) | −0.0725*** (0.0405) | 0.0446 (0.0460) |
| γ | 0.0526* (0.0161) | 0.0161* (0.0162) | 0.0613* (0.0165) | 0.5190 (0.0944) | 0.6226* (0.0862) |

| Variance equation |       |       |       |       |       |       |
| Coefficient |       |       |       |       |       |       |
| ω | 1.0044* (0.0405) | 0.4196* (0.0169) | 0.7671* (0.0560) | 1.2381* (0.0561) | 1.7419* (0.0668) | 0.4334* (0.0271) |
| α | 0.2743* (0.0146) | 0.1309* (0.0079) | 0.1456* (0.0102) | 0.2648* (0.0160) | 0.3596* (0.0179) | 0.0958* (0.0058) |
| γ | 0.0005 (0.0192) | −0.0225* (0.0087) | −0.0085 (0.0118) | −0.0488* (0.0172) | −0.0979* (0.0230) | −0.0067 (0.0072) |
| β | 0.6497* (0.0103) | 0.8462* (0.0048) | 0.7842* (0.0114) | 0.6782* (0.0108) | 0.6059* (0.0106) | 0.8764* (0.0044) |

Coefficients refer to the estimates of the mean and variance equations of the following ARMA-GARCH(1,1) model: Mean Equation: \( R_t = α_t + \theta R_{t-1} + \gamma E_{R_t} + \epsilon_t \); Variance Equation: \( \log(h_t) = \omega + \alpha \frac{\epsilon_{t-1}}{\sqrt{h_{t-1}}} + \gamma \frac{\epsilon_{t-1}}{\sqrt{h_{t-1}}} + \beta \log(h_{t-1}) \); Numbers in parentheses indicate standard errors. *, **, *** indicate the significance level at 1%, 5% and 10% respectively.

Table 6: The EGARCH representation of tourism stock returns model

| Coefficient | AYCES | MAALT | MARTI | METUR | PKENT | TEKTU |
|-------------|-------|-------|-------|-------|-------|-------|
| Mean equation |       |       |       |       |       |       |
| α | 0.0631** (0.0312) | 0.0366 (0.0354) | 0.0043 (0.0347) | −0.0199 (0.0359) | −0.0737 (0.0267) | 0.0257 (0.0443) |
| β | 0.1089*** (0.0382) | 0.0718*** (0.0424) | −0.1096* (0.0390) | 0.0816 (0.0352) | −0.0549*** (0.0336) | 0.0410 (0.0451) |
| γ | 0.0379*** (0.0151) | −0.0676* (0.0158) | −0.0403 (0.0160) | 0.5040* (0.0681) | −0.6237* (0.0591) |

| Variance equation |       |       |       |       |       |       |
| Coefficient |       |       |       |       |       |       |
| ω | 0.0002 (0.0068) | 0.0444* (0.0056) | 0.0024 (0.0103) | 0.0136 (0.0101) | 0.0862* (0.0141) | 0.0348* (0.0072) |
| α | 0.4273* (0.0124) | 0.3073* (0.0104) | 0.2769* (0.0117) | 0.3958* (0.0135) | 0.4869* (0.0137) | 0.2093* (0.0077) |
| γ | 0.0169*** (0.0087) | 0.0259* (0.0064) | 0.0170* (0.0068) | 0.0324* (0.0080) | 0.0516* (0.0098) | 0.0070 (0.0051) |
| β | 0.8537* (0.0054) | 0.9237* (0.0029) | 0.9080* (0.0062) | 0.8740* (0.0054) | 0.8278* (0.0071) | 0.0954* (0.0030) |

Coefficients refer to the estimates of the mean and variance equations of the following ARMA-GARCH1(1,1) model: Mean Equation: \( R_t = α_t + \alpha R_{t-1} + \gamma E_{R_t} + \epsilon_t \); Variance Equation: \( \log(h_t) = \omega + \alpha \frac{\epsilon_{t-1}}{\sqrt{h_{t-1}}} + \gamma \frac{\epsilon_{t-1}}{\sqrt{h_{t-1}}} + \beta \log(h_{t-1}) \); Numbers in parentheses indicate standard errors. *, **, *** indicate the significance level at 1%, 5% and 10% respectively.
show that the coefficient that represents the time-independent component of volatility ($\omega$) is small in magnitude, positive, and statistically significant in only three cases. However, the ARCH and GARCH components, a $\alpha$ and $\beta$ respectively, are large and significant in all cases. These findings imply that the time-independent components of volatility are more important than the time-independent component of volatility. Second, the significant and positive ARCH coefficient, $\alpha$, imply the presence of volatility clustering and of the tendency of shocks to persist. The GARCH coefficients ($\beta$’s) are positive, significant and $<1$ for tourism firms. This suggests that the volatility in tourism stock returns is very persistent, meaning that the volatility remains high and will dissipate very slowly over time. The results in Table 6 shows that the coefficients, $\gamma$, which represent the asymmetric and leverage effects of shocks to current volatility, are positive and significant in five out of six cases. The positive and significant $\gamma$ coefficient suggests that there is no leverage effect and that only asymmetric effects exist, implying that the impact of positive news on current volatility is larger than negative news of the same magnitude.

It should also be noted that the statistical inferences carried out above about the tourism stock return models are valid inferences since the stock return series subject to empirical analysis satisfy the non-negativity, stability and stationarity conditions. As seen from Tables 4-6, the stationarity conditions of $\alpha+\beta<1$ and $|\beta|<1$ for the EGARCH model, and a non-negativity condition of $\alpha>0$ and $\beta>0$ for the GARCH and GJR models, are satisfied. These sufficient conditions make sure that the Quasi-Maximum Likelihood Estimators (QMLE) are consistent and asymptotically normal (Chang et al., 2014; McAller et al., 2007).

5. CONCLUSION

This study examined the impact of exchange rate risk on stock returns for Turkish tourism firms listed in Borsa Istanbul. To cope with the time varying properties of volatility, the ARMA model of tourism stock returns are estimated with the GARCH, GJR (or TGARCH), and EARCH variance specifications. The estimation results showed that the EGARCH model delivered valuable information on the impact of exchange rate risk on stock returns of tourism firms by handling volatility in stock return series properly. The results indicate that the effect of an increase in exchange rate risk differs among tourism firms significantly. While the impact of exchange rate risk on stock returns of tourism firms is positive and significant for 3 cases, it is negative and significant for 2 firms.

The empirical findings on the volatility of stock returns obtained from the variance equation indicate that the time-dependent components of volatility are clearly more important than the time-independent component of volatility in predicting current volatility in all models. Furthermore, the results suggest that the volatility of stock returns is highly persistent, and the volatility at time $t$ is more sensitive to past period volatility than past surprises in the market. Surprisingly, the sign of the $\gamma$ coefficient, which shows the asymmetry of shocks on volatility, is negative in the GJR model and positive in the EGARCH model suggesting that the impact of negative news do not outweigh positive news or that the impact of positive news on volatility is higher than the impact of negative news in the market. This suggests that while positive innovations like a market boom rise volatility of returns, negative innovations like market stagnation leads to a decline in volatility. This rather surprising finding requires further analysis of stock returns of Turkish tourism firms.

REFERENCES

Bollerslev, T. (1986), Generalized autoregressive conditional heteroskedasticity. Journal of Econometrics, 31, 307-327.
Çelik, I.E. (2019), Assessing the impact of bank risk factors on Turkish bank’s stock returns using the egarch-m model. Üçüncü Sektör Sosyal Ekonomi Dergisi, 54(2), 811-827.
Chang, C.L., Hsu, H.K., McAller, M. (2013), Is small beautiful? Size effects of volatility spillovers for firm performance and exchange rates in tourism. The North American Journal of Economics and Finance, 26, 519-534.
Chang, C.L., Hsu, H.K., McAller, M. (2014), The impact of China on stock returns and volatility in the Taiwan tourism industry. North American Journal of Economics and Finance, 29, 381-401.
Cordella, T., Rojas, A.O. (2017), Financial Globalization and Market Volatility an Empirical Appraisal, World Bank Group, Policy Research Working Paper 8091. Washington, D.C: World Bank Group.
Dickey, D., Fuller, W. (1981), Likelihood ratio statistics for autoregressive time series with a unit root. Econometrica, 49, 1057-1072.
Doğanak, H., Özmen, M., Yücel, E. (2010), IMKB’de sektörler açığa döviz kuru duyarlılığının incelenmesi. Çukurova Üniversitesi Sosyal Bilimler Enstitüsü Dergisi, 19(3), 63-86.
Dornbusch, R., Stanley, F. (1980), Exchange rates and the current account. American Economic Review, 70, 960-971.
Elyasiani, E., Mansur, I. (1998), Sensitivity of bank stock returns distribution to changes in the level of volatility of interest rate: A GARCH-M model. Journal of Banking and Finance, 22, 535-563.
Enders, W. (2015), Applied Econometrics Time Series. United States: John Wiley & Sons Inc.
Engle, R.F., Granger, C.W.J. (1987), Cointegration and error correction: Representation, estimation and testing. Econometrica, 50, 987-1007.
Engle, R.F., Ng, V., Rothschild, M., (1990), Asset pricing with a factor-ARCH covariance structure: Empirical estimates for treasury bills. Journal of Econometrics, 45, 213-238.
Eroğan, M. (2018), Turizm işletmelerinin finansal performanslarının etkileyen firmaya özgü, sektörü ve makroekonomik belirleyicileri. Mustafa Kemal Üniversitesi Sosyal Bilimler Enstitüsü Dergisi, 15(41), 101-127.
Ergül, N. (2014), BIST-turizm sektöründeki şirketlerin finansal performans analizi. Çankırı Karatekin Üniversitesi Iktisadi ve İdari Bilimler Fakültesi Dergisi, 4(1), 325-340.
Fauziah, F., Moeljadi, M., Ratnawati, K. (2015), Dynamic relationship between exchange rates and stock prices in Asia, 2009-2013. Journal of Economics Finance and Accounting, 2(1), 124-134.
Flannery, M.J., Hameed, A.S., Harjes, R.H. (1997), Asset pricing, time-varying risk premia and interest rate risk. Journal of Banking and Finance, 21, 315-335.
Glosten, L., Ravi, J., Runkle, D. (1993), On the relation between the expected value and the volatility of the nominal excess return on stocks. Journal of Finance, 48, 1779-1801.
Gökmenoglu, K.K., Hadood, A.A.A. (2019), Volatility spillovers and time varying correlation for Chinese tourism firms. Asia Pacific Journal of Tourism Research, 24(6), 584-596.
Güdük, T. (2018), The relationship between tourism revenues and financial ratios of enterprises in the BIST tourism: Panel data analysis. Prizren Social Science Journal, 2(2), 35-42.
Hsiao, C.M. (2017), Enterprise risk management with foreign exchange exposures: Evidence from Taiwan tourism industry. Asian Economic and Financial Review, 7(9), 822-906.

Hyde, S. (2007), The response of industry stock returns to market, exchange rate and interest rate risks. Managerial Finance, 33, 693-709.

Kandil Göker, İ.E., Uysal, B. (2020), Turizm sektöründe faiz oranı ve döviz kursunun karlılık üzerindeki etkisi. Ömer Halisdemir Üniversitesi İktisadi ve İdari Bilimler Fakültesi Dergisi, 13(1), 159-170.

Karadeniz, E., İskenderoğlu, Ö. (2011), İstanbul menkul kıymetler borsası’nda işlem gören turizm işletmelerinin aktif karlılığını etkileyen değişkenlerin analizi. Anatolia: Turizm Araştırmaları Dergisi, 22(1), 65-75.

Kasman, S., Vardar, G., Tunç, G. (2011), The impact of interest rate and exchange rate volatility on banks’ stock returns and volatility: Evidence from Turkey. Economic Modelling, 28, 1328-1334.

Kose, M.A., Prasad, E.S., Terrones, M.E. (2009), Does financial globalization promote risk sharing? Journal of Development Economics, 89, 258-270.

Kutlu, M., Karakaya, A. (2019), Borsa istanbul tourism index volatility: Markov regime switching arch model, special issue on applied economics and finance. Journal of Yasar University, 14, 18-24.

Lee, S.K., Jang, S. (2010), Internationalization and exposure to foreign currency risk: An examination of lodging firms. International Journal of Hospitality Management, 29(4), 701-710.

Lin, C.H. (2012), The comovement between exchange rates and stock prices in the Asian emerging markets. International Review of Economics and Finance, 22, 161-172.

Liu, H.C., Hung, J.C. (2010), Forecasting S&P-100 stock index volatility: The role of volatility asymmetry and distribution assumption in GARCH models. Expert Systems with Applications, 37(7), 4928-4934.

Mansur, I., Elyasiani, E., (1995), Sensitivity of bank equity returns to the level and volatility of interest rates. Managerial Finance, 21, 58-77.

McAleer, M., Chan, F., Marinova, D. (2007), An econometric analysis of asymmetric volatility: Theory and application to patents. Journal of Econometrics, 139(2), 259-284.

Mohapatra, S.M. (2017), A comparison of exchange rate exposure between manufacturing vis-à-vis service sector firms in India. Economic Papers A Journal of Applied Economics and Policy, 36(1), 75-85.

Nelson, D.B. (1991), Conditional heteroskedasticity in asset returns: A new approach. Econometrica, 37-470.

Oby, P., Sil, S., Abuizam, R. (2015), Tourism stock, implied volatility and hedging: A vector error correction study. Journal of Accounting and Finance, 15(8), 30-39.

Olugbode, M., El-Masry, A., Pointon, J. (2014), Exchange rate and interest rate exposure of UK industries using first-order autoregressive exponential GARCH-in-mean (EGARCH-M) approach, The Manchester School, 82(4), 409-464.

Özçelik, H., Kandemir, B. (2015), BIST’dede işlem gören turizm işletmelerinin toplas yöntemi ile finansal performanslarının değerlendirilmesi. Balıkesir University The Journal of Social Sciences Institute, 18(33), 98-114.

Phillips, P.C.B., Perron, P., (1988), Testing for a unit root in time series regression. Biometrika, 75, 335-346.

Saunders, A., Yourougou, P., (1990), Are banks special? The separation of banking from commerce and interest rate risk. Journal of Economics and Business, 42, 171-182.

Sehgal, S., Agrawal, T.J., (2017), Bank risk factors and changing risk exposures in the pre-and post-financial crisis periods: An empirical study for India. Management and Labour Studies, 42(4), 356-378.

Song, F., (1994), A two factor ARCH model for deposit-institution stock returns. Journal of Money Credit and Banking, 26, 323-340.

Soyaslan, E., (2019), Döviz kuru ile BIST turizm endeksi arasındaki ilişkinin analizi. Uluslararası Toplum Araştırmaları Dergisi, 12(18), 772-793.

Stiglitz, J.E. (2004), Capital-market liberalization, globalization, and the IMF. Oxford Review of Economic Policy, 20, 57-71.

Umultu, M., Akdeniz, L., Altay, S. (2010), The degree of financial liberalization and aggregated stock-return volatility in emerging markets. Journal of Banking and Finance, 34, 509-521.

Vardar, G., Aksoy, G., Can, E. (2008), effects of interest rate and exchange rate on volatility and return of sector price indices at Istanbul stock exchange. European Journal of Economics Finance and Administrative Sciences, 11, 26-35.

Yamak, N., Koc, F., Köyel, F., (2018), Asymmetric relationship between exchange rate volatility and stock market index volatility. BJSS Balkan Journal of Social Sciences, 7(14), 171-187.