What accounts for the association between late preterm births and risk of asthma?

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ABSTRACT

Background: Although results of many studies have indicated an increased risk of asthma in former late preterm (LPT) infants, most of these studies did not fully address covariate imbalance.

Objective: To compare the cumulative frequency of asthma in a population-based cohort of former LPT infants to that of matched term infants in their early childhood, when accounting for covariate imbalance.

Methods: From a population-based birth cohort of children born 2002–2006 in Olmsted County, Minnesota, we assessed a random sample of LPT (34 to 36 6/7 weeks) and frequency-matched term (37 to 40 6/7 weeks) infants. The subjects were followed-up through 2010 or censored based on the last date of contact, with the asthma status based on predetermined criteria. The Kaplan-Meier method was used to estimate the cumulative incidence of asthma during the study period. Cox models were used to estimate the hazard ratio and 95% confidence interval for the risk of asthma, when adjusting for potential confounders.

Results: LPT infants (n = 282) had a higher cumulative frequency of asthma than did term infants (n = 297), 29.9 versus 19.5%, respectively; p = 0.01. After adjusting for covariates associated with the risk of asthma, an LPT birth was not associated with a risk of asthma, whereas maternal smoking during pregnancy was associated with a risk of asthma.

Conclusion: LPT birth was not independently associated with a risk of asthma and other atopic conditions. Clinicians should make an effort to reduce exposure to smoking during pregnancy as a modifiable risk factor for asthma.

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identified as small for gestational age or large for gestational age based on birth weight cutoffs for sex and completed week of gestation. Among LPT infants who were appropriate for gestational age, 200 subjects were randomly selected. Term subjects were frequency matched based on sex, birth hospital, size for gestational age, gestational age distribution, and year of birth.

Asthma Ascertainment
The primary outcome was the development of asthma as defined by criteria previously described and noted in Supplemental Table 2. Definite and probable asthma cases were considered to be subjects with asthma because most probable asthmas become definite asthmatics over time. Secondary outcomes included physician-diagnosed other atopic conditions (eczema and/or atopic dermatitis, and allergic rhinitis and/or hay fever).

Other Variables
We reviewed the entire medical record, including the birth certificate, to collect all pertinent covariates and potential confounders for the study (Table 1).

Statistical Analysis
Descriptive statistics (median [25th, 75th percentile] for continuous variables, and counts [percentages]) were used to compare the two cohorts. For each characteristic, the difference was compared by using logistic regression models when adjusting for matching factors. The cumulative incidence of asthma during the first 8 years of life was calculated by using Kaplan-Meier estimates and the log-rank test for statistical significance. The risk of asthma between LPT and term infants was compared by using unadjusted Cox proportional hazards models. Characteristics were tested for association with LPT by using logistic regression models. A multivariable Cox regression model was created to assess the independent association for the risk of asthma univariately by using Cox proportional hazards models. A multivariable Cox regression model was created to assess the relationship between LPT and the risk of asthma after controlling for all potential confounders (Table 1). Similar analyses were conducted for risk of other atopic conditions. All statistical analyses were adjusted for matching factors and performed by using SAS software, version 9.3 (SAS Institute Inc., Cary, NC).

RESULTS
During 2002–2006, there were 8113 eligible births in Olmsted County. Eleven infants were excluded because there was no identifiable birth weight. Of the remaining 8102 infants, 647 were LPT and 7455 were term births. In the LPT cohort, we enrolled 29 small for gestational age, a random sample of 195 appropriate for gestational age, and 58 large for gestational age infants. The final term cohort included the following: 27 small for gestational age, 200 appropriate for gestational age, and 70 large for gestational age infants. The LPT cohort was slightly smaller than the term cohort because not all excluded subjects had a suitable replacement match. The demographic and clinical characteristics are shown in Table 1. There were significant differences in risk factors for asthma, such as smoking during pregnancy, passive smoke exposure after birth, and a family history of asthma. Each of these factors was more common in the LPT cohort, whereas breast milk exposure was less common in the LPT group (78 versus 85%; p = 0.04).

The cumulative incidence of asthma is shown in Fig. 1, with each “step” denoting the age in years in which a subject met the criteria for asthma. The cumulative incidence of asthma at 5 years in the LPT cohort was 29.9% compared with 19.5% in the term cohort (unadjusted hazard ratio [HR] 1.67 [95% confidence interval [CI], 1.13–2.45]; p = 0.01) (Table 1). We analyzed the potential confounders related to LPT in univariate Cox models as summarized in Table 1. The following factors were associated with an increased risk of asthma: maternal smoking during pregnancy (HR 3.05 [95% CI, 1.86–5.01]), passive smoke exposure after birth (HR 1.99 [95% CI, 1.29–3.07]), family history of asthma (HR 3.26 [95% CI, 2.22–4.78]), and family history of other atopic conditions (HR 1.30 [95% CI, 0.89–1.92]). Also, after controlling for all pertinent covariates, we found that the association between LPT births and the risk of asthma was significantly attenuated and lacked statistical significance (adjusted HR 1.44; p = 0.07) (Table 2). However, a family history of asthma and maternal smoking during pregnancy imparted an independent association for the risk of asthma (Table 2). LPT births were not associated with a risk of other atopic conditions (HR 1.26 [95% CI, 0.93–1.71]; p = 0.13).

DISCUSSION
We found that healthy LPT infants had a higher incidence of asthma by school age, with an HR of 1.67 (p = 0.01) in an unadjusted Cox proportional hazard model. However, after controlling for maternal smoking during pregnancy, passive smoke exposure after birth, and a family history of asthma and other atopic conditions, the HR for asthma in LPT infants was significantly attenuated, at 1.44, and was no longer significantly different (p = 0.07). A recent review of outcomes of LPT birth specifically addresses what is known on respiratory outcomes in this population.
strate contradictory findings, some of which may be attributed to subject bias and difficulty controlling confounding variables.8,10,16 We also recently demonstrated that a historical birth cohort of LPT infants had a significantly increased risk of asthma compared with term infants.2 After fully adjusting for covariate imbalance by using propensity score matching, there was no significant difference in the incidence of asthma (HR 1.13; \(p = 0.56\)).2

The key finding of our study was the role of maternal smoking during pregnancy as a modifiable environmental exposure, which accounted for the association between LPT and the risk of asthma. Several other studies noted the importance of maternal smoking during pregnancy and the impact on the risk of asthma.17–21 Smoking during pregnancy has been shown to increase the risk of preterm birth, which likely accounts for the difference of this confounder between the two cohorts in our study.22,23 This confounder must be included in any study that assesses the association of preterm birth and asthma.

Our study had inherent limitations as a retrospective study design. Our study population was predomi-

| Variables                        | N   | Events, no. (%) | HR (95% CI) | p Value |
|----------------------------------|-----|-----------------|-------------|---------|
| Birth                            |     |                 |             |         |
| Term                             | 297 | 43 (14)         | Reference   | 0.01    |
| LPT                              | 282 | 65 (23)         | 1.67 (1.13–2.45) |         |
| Length of hospital stay          | 574 |                 | 1.02 (0.97–1.06) | 0.41    |
| White                            |     |                 |             |         |
| No                               | 166 | 31 (19)         | Reference   | 0.74    |
| Yes                              | 413 | 77 (19)         | 0.93 (0.61–1.42) |         |
| Cesarean section                 |     |                 |             | 0.69    |
| No                               | 417 | 75 (18)         | Reference   |         |
| Yes                              | 162 | 33 (20)         | 1.09 (0.71–1.67) |         |
| GBS                              |     |                 |             | 0.69    |
| Negative                         | 321 | 57 (18)         | Reference   |         |
| Positive                         | 113 | 20 (18)         | 1.05 (0.62–1.75) |         |
| Unknown                          | 145 | 31 (21)         | 1.22 (0.78–1.91) |         |
| Apgar score at 1 minute          |     |                 |             | 0.32    |
| 0–3                              | 4   | 2 (50)          | Reference   |         |
| 4–6                              | 59  | 14 (24)         | 0.52 (0.11–2.37) |         |
| >6                               | 516 | 92 (18)         | 0.38 (0.09–1.62) |         |
| Multiple gestations              |     |                 |             | 0.73    |
| No                               | 512 | 94 (18)         | Reference   |         |
| Yes                              | 67  | 14 (21)         | 1.11 (0.62–2.00) |         |
| Breastfeeding history            |     |                 |             | 0.97    |
| No                               | 107 | 21 (20)         | Reference   |         |
| Yes                              | 461 | 86 (19)         | 1.01 (0.62–1.65) |         |
| Maternal smoking during pregnancy|     |                 |             | <0.001  |
| No                               | 467 | 74 (16)         | Reference   |         |
| Yes                              | 53  | 21 (40)         | 3.05 (1.86–5.01) |         |
| Unknown (missing)                | 59  | 13 (22)         | 1.54 (0.78–3.06) |         |
| Passive smoking exposure after birth |  |                 |             | 0.01    |
| Passive                          | 106 | 32 (30)         | Reference   |         |
| No exposure                      | 409 | 67 (16)         | 0.50 (0.330.77) |         |
| Unknown (missing)                | 64  | 9 (14)          | 0.56 (0.26–1.18) |         |
| Family history of asthma         |     |                 |             | <0.001  |
| No                               | 447 | 58 (13)         | Reference   |         |
| Yes                              | 132 | 50 (38)         | 3.26 (2.22–4.78) |         |
| Family history of other atopic conditions | |                 |             | 0.18    |
| No                               | 371 | 63 (17)         | Reference   |         |
| Yes                              | 208 | 45 (22)         | 1.30 (0.89–1.92) |         |

LPT = Late preterm; HR = hazard ratio; CI = confidence interval; GBS = group B Streptococcus.
nantly suburban and white, so the findings may not be generalizable to other populations and settings. Because information regarding smoking during pregnancy and passive smoke exposure after birth was collected from the medical record, we were unable to accurately assess the precise timing of maternal smoking during pregnancy. Future prospective studies that focus on the timing of smoke exposure in utero and passive exposure after birth may help to further define the full impact of these confounders. Our study had several strengths. First, we minimized selection bias by sampling a population-based birth cohort. Second, our use of standard definitions of LPT and utilization of birth records to determine subjects’ gestational ages.

Figure 1. Cumulative incidence of asthma in late preterm and term infants, and the parameter estimate by univariate analysis.

Table 2 Multivariable Cox models for the association between LPT and the risk of asthma and other atopic conditions

| Variables                                    | N   | HR (95% CI)         | p Value |
|----------------------------------------------|-----|---------------------|---------|
| Birth                                        |     |                     |         |
| Term                                         | 297 | Reference           | 0.07    |
| LPT                                          | 282 | 1.44 (0.97–2.15)    |         |
| Maternal smoking during pregnancy            |     |                     |         |
| No                                           | 467 | Reference           | 0.02    |
| Yes                                          | 53  | 2.24 (1.20–4.18)    |         |
| Unknown (missing)                            | 59  | 1.98 (0.94–4.15)    |         |
| Passive smoking exposure after birth         |     |                     | 0.96    |
| No exposure                                  | 106 | Reference           |         |
| Yes                                          | 409 | 0.94 (0.54–1.65)    |         |
| Unknown (missing)                            | 64  | 1.04 (0.46–2.35)    |         |
| Family history of asthma                     |     |                     | <0.001  |
| No                                           | 447 | Reference           |         |
| Yes                                          | 132 | 3.24 (2.10–5.02)    |         |
| Family history of other atopic condition     |     |                     | 0.56    |
| No                                           | 371 | Reference           |         |
| Yes                                          | 208 | 0.88 (0.58–1.35)    |         |

LPT = Late preterm; HR = hazard ratio; CI = confidence interval.
minimized the risk of misclassification bias of LPT births. Third, we improved the accuracy of the asthma diagnosis by using predetermined criteria instead of International Classification of Diseases codes or self-report for asthma. Fourth, the main findings and the identified confounders were consistent between our current and previous study; similarly, the absence of an association between LPT birth and the risks of other atopic conditions demonstrated consistency and biologic coherence.

CONCLUSION

LPT birth is not independently associated with the risk of asthma and other atopic conditions. We found a critical role of smoking during pregnancy in the development of asthma among LPT infants. Smoking cessation programs during pregnancy have not been highly successful in the past, and more efforts could be made in this area because it would likely reduce the risk of preterm births and asthma. Although clinicians should make an effort to reduce exposure to smoking during pregnancy, parents whose child was born LPT should be reassured about their child’s own risk for eventually developing asthma.

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