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Impacts of local adaptation of forest trees on associations with herbivorous insects: implications for adaptive forest management

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Keywords
adaptive forest management, climate matching, gallwasp, local adaptation plant–insect interactions, population nonindependence, provenance trials, Quercus petraea.

Abstract
Disruption of species interactions is a key issue in climate change biology. Interactions involving forest trees may be particularly vulnerable due to evolutionary rate limitations imposed by long generation times. One mitigation strategy for such impacts is Climate matching – the augmentation of local native tree populations by input from nonlocal populations currently experiencing predicted future climates. This strategy is controversial because of potential cascading impacts on locally adapted animal communities. We explored these impacts using abundance data for local native gallwasp herbivores sampled from 20 provenances of sessile oak (Quercus petraea) planted in a common garden trial. We hypothesized that non-native provenances would show (i) declining growth performance with increasing distance between provenance origin and trial site, and (ii) phenological differences to local oaks that increased with latitudinal differences between origin and trial site. Under a local adaptation hypothesis, we predicted declining gallwasp abundance with increasing phenological mismatch between native and climate-matched trees. Both hypotheses for oaks were supported. Provenance explained significant variation in gallwasp abundance, but no gall type showed the relationship between abundance and phenological mismatch predicted by a local adaptation hypothesis. Our results show that climate matching would have complex and variable impacts on oak gall communities.

Introduction
Ecosystems across the world are experiencing profound climate change (IPCC, 2013), driving population responses whose effects can influence community structure and function (Walther et al. 2002; Parmesan and Yohe 2003). Forest ecosystems are expected to be particularly sensitive to pressures imposed by climate because the long generation times of trees result in low rates of adaptation (Lindner et al. 2010). Furthermore, because forest trees are foundation (and also often keystone) species for associated communities of plants, animals, fungi and microbes (Kennedy and Southwood 1984; McEwan et al. 2011; Lindenmayer et al. 2014), effects on them are expected to cascade through associated trophic levels (Frelch et al. 2012). Consequently, there is considerable interest in developing adaptive forest management strategies to preserve ecosystem productivity and services in the face of climate change (Spittlehouse and Stewart 2003; Broadmeadow et al. 2005; Aitken et al. 2008; Bower and Aitken 2008; Bolte et al. 2009). Assisted migration, in the form of climate matching, has emerged as one potential strategy facilitating forest adaptation to predicted future climates (Hoegh-Guldberg et al. 2008; Dawson et al. 2011), particularly in commercial
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for in situ adaptation may be limited (Broadmeadow et al. 2005; but see Cavers and Cottrell 2014). Climate matching is based on the premise that natural forest stands are locally adapted to current local climates (Savolainen et al. 2007, 2013; Alberto et al. 2013a), and for a given species and planting location involves the following steps: (i) identify climate variables that predict performance for the target tree species across its range; (ii) use regional climate models to predict future climates at the proposed planting site; (iii) identify locations within the species range that currently experience the predicted conditions; (iv) plant seed from tree populations in the identified locations (termed provenances in forestry) in the expectation that resulting trees will be well matched to future climates, and contribute adaptive variation that accelerates subsequent local adaptation (Broadmeadow et al. 2005).

As remediation strategies, climate matching and other forms of assisted migration remain subjects of vigorous debate (Dawson et al. 2011; Lunt et al. 2013; Cavers and Cottrell 2014). Forests are biologically diverse ecosystems, and the value of climate matching, or any other strategy that modifies the genetic, physical or spatial structure of tree populations, must be considered not only in terms of its effectiveness in promoting forest adaptation and resilience, but also its impacts on local associated biodiversity. Community genetic research has shown that the organisms associated with forest stands reflect the genotypes and genetic diversity of the host trees (Dungey et al. 2000; Wimp et al. 2005; Whitham et al. 2012), and a range of studies have shown adaptation of plants to local herbivores and vice versa (Sork et al. 1993; Hanks and Denno 1994; Egan and Ott 2007; Tack and Roslin 2010; Bernhardsson et al. 2013). As such, the introduction of new tree genotypes as part of a climate matching strategy may drive changes in the identity, abundance and genetics of associated species (Lefevre et al. 2014). However, it is difficult to forecast the direction and magnitude of such changes. While some studies predict increased abundance of herbivores on non-native tree provenances (Bernhardsson et al. 2013), others suggest the opposite through local herbivore adaptation to plant phenology (Van Asch and Visser 2007; Visser 2008; Pearse and Karban 2013). The key issues are the extent to which host plant provenances vary in phenotypic traits relevant to herbivore success, and the magnitude of herbivore population responses to such variation.

Here we explore the community effects of climate matching by quantifying the impact of provenance on (i) the phenotype (growth performance and phenology) of an important and widespread European oak, Quercus petraea, and (ii) the abundance of associated insect herbivores. Our approach exploits an established INRA forestry trial at La Petite Charnie in Northwest France, where 96 provenances of Q. petraea grow surrounded by natural oak forest and are exposed to native herbivores. Our study focuses on oak cynipid gallwasps (Hymenoptera; Cynipidae), specialized herbivores whose larvae develop in galls on their oak hosts (Schönrogge et al. 2000; Stone and Schönrogge 2003; Harper et al. 2004). Gall induction requires access to host oak tissues at the correct developmental stage (Harper et al. 2004), and oak cynipids are highly sensitive to host plant phenology (Askew 1962; Crawley and Akhteruzzaman 1988). These aspects of gallwasp biology, coupled with a partly parthenogenetic lifecycle, are associated with local adaptation to specific host plant populations (Mopper et al. 2000; Mopper 2005; Egan and Ott 2007; Tack et al. 2010) and individuals (Egan and Ott 2007) and make galls an appropriate sentinel taxon for exploration of the potential effects of climate matching.

We expect the performance of oak provenances to correlate positively with the extent to which they are adapted to prevailing climatic conditions (Broadmeadow et al. 2005; Lefèvre et al. 2014). This hypothesis predicts that performance of locally native provenances will decline over time – and recent rapid climate change may mean that populations of long-lived trees are already somewhat mismatched to the conditions in which they grow. The hypothesis also predicts that provenances selected to match conditions furthest in the future should be least well matched to current conditions and so perform least well in the present. We test this hypothesis using multiple performance measures for 20 provenances planted at the La Petite Charnie provenance trial, selected to maximize the available range in source latitude and longitude as well as the mismatch in climate between the provenance trial and the provenance origin. We then test two contrasting hypotheses linking herbivorous insect abundance to observed variation in performance among oak provenances. The plant stress hypothesis predicts that herbivores should show improved performance and abundance on physiologically stressed plants, due to higher levels of mobilized nitrogen in stressed plant tissues and an associated increase in resource quality for herbivores (White 1969, 1974, 2009). In contrast, the plant vigour hypothesis (Price 1991; Cornelissen et al. 2008) predicts increased herbivore performance and abundance on more vigorous plants or plant organs (those with high growth rates or large ultimate size relative to the population mean) through elevated nutrient content and reduced concentrations of defensive compounds. Here we test both hypotheses using abundance data across provenances for 20 cynipid gall types.

Phenological matching of key biological events is crucial for many interactions between forest trees and associated organisms. For example, synchronization of breeding with high plant resource availability (budburst for leaf-eating herbivores) is central to reproductive success in herbivores and associated trophic levels. We thus expect associated
organisms to show local adaptation to host plant phenology (Ducousso et al. 1996; Yukawa 2000; Van Asch and Visser 2007; Van Asch et al. 2012; Alberto et al. 2013b). Many plants show latitudinal clines in phenological events, such as the timing of budburst, whose onset is often determined by photoperiodic cues (Savolainen et al. 2007; Atten et al. 2008) or the interaction between photoperiod and temperature (Laube et al. 2014). This is relevant to climate matching because, in a warming world, it often involves introducing plants to higher latitudes from lower latitudes. As a result, phenological mismatches with local herbivores might occur by disrupting any adaptation to the local combination of thermal and photoperiod cues, with associated impacts on the fitness of herbivores and trophically linked taxa such as natural enemies (Visser 2008; Thomson et al. 2010; Kerstes and De Jong 2011; Van Asch et al. 2012). Here we quantify the phenotypic mismatch in budburst date between the same 20 provenances and assess the impact of phenological mismatch on cynipid gall abundance. Specifically, we test the local adaptation hypothesis, that is that herbivore abundance will decline with increasing phenological difference between introduced and locally native oak provenances. We distinguish phenological effects from other potential differences between provenances (e.g., variation in chemical defences) by incorporating provenance phenology into our analyses.

Our analytical approach recognizes that tree populations are linked by shared common ancestry and gene flow and hence are unlikely to represent statistically independent replicates (Stone et al. 2011). Here we investigate the impact of potential nonindependence by comparing the results of analyses that incorporate the expected variance–covariance among populations in a generalized linear mixed model framework (Hadfield and Nakagawa 2010).

Materials and methods
Provenance trial design
Provenance effects were explored using a large experimental trial of *Q. petraea* at La Petite Charnie in Northwest France (Fig. 1; 48.086° north, 0.168° west; details on the planting and layout of the trial are provided in Data S1). We selected 20 provenances whose origins span 15° of latitude and 34° of longitude, incorporating the greatest available geographic separation and source environmental variation (Fig. 1, Table S1 in Data S1), and included the closest provenance to the trial site (Forêt de Bercé, approximately 50 km from La Petite Charnie). Each sampled provenance is planted in multiple replicate plots of 24 trees distributed across five soil zones (see Data S1). We selected two plots per soil zone per provenance, giving a total of 200 study plots containing 4800 trees. Previous INRA surveys provided data for each tree for the following

Figure 1 Location of the Petite Charnie provenance trials and the 20 selected study provenances. Sites are identified by number in Table 1. Countries contributing provenances are indicated by their two letter ISO codes.
phenotypic traits (see Data S1 for illustrations of character states):

1 **Budburst** – spring budburst phenology was scored on a 0–5 ordinal scale, where 0 is a dormant bud and 5 is a fully open bud, in spring 1995 (this closely parallels that used in other studies on oak (Crawley and Akhteruzzaman 1988; Tack et al. 2010)). A higher score on the survey date indicates earlier budburst. Although oak phenology is not known to show strong $G \times E$ interactions across years (Alberto et al. 2011), we nevertheless repeated phenology scoring in 2009 for 582 trees in a subset of eight provenances. Phenology scores for the same tree in 1995 and 2009 were significantly and positively correlated (see Data S4). We thus regard using the 1995 data for all 20 provenances as unlikely to produce misleading or contrary results.

2 **Diameter at breast height (DBH)** – tree diameter at a height of 1.3 m, measured during winter 2001–2002. A higher value is associated with more rapid growth and higher plant vigour.

3 **Form** – a measure based on tree shape, recorded on an ordinal scale from 1 to 10 during winter 2001–2002. Stressed trees with bushy, irregular growth and poor apical dominance receive a low score, while healthy trees with straight stems, even branching and strong apical dominance score highly. These phenotypic measures showed substantial variation among provenances (discussed in more detail below), providing the basis for examination of impacts of phenological matching (Budburst) and performance (DBH, Form) on herbivore abundance (Table 1, Fig. 2).

**Herbivore sampling**

The cynipid species surveyed in this study all have broad geographic distributions; like *Q. petraea* and closely related oaks, they span the Western Palearctic from the Iberian Peninsula eastwards to Iran and the Caucasus. Cynipid gall abundance was sampled across the 20 selected provenances in 2008 and 2009. Gallwasps have a spring (sexual) and autumn (asexual) gall generation each year, and both were sampled, giving four surveys in total. For each of the 200 study plots, we sampled galls from 12 of the 24 available trees, a total of 2400 trees. To minimize edge effects, surveyed trees comprised the two internal rows of six trees within each plot. During each survey, using a pole pruner giving us approximately 4 m reach we selected 10 twigs on each tree, each consisting of the last 2 years’ growth identified using ring scars on the bark. For each twig, we surveyed the galls on all parts (buds, leaves, stem, catkins, acorns; there are no cynipid galls on female flowers in Europe) of its terminal shoot (a module of woody growth from Table 1).

Table 1. Summary of the 20 studied provenances of *Quercus petraea*, ordered by longitude from west to east showing their three digit INRA provenance codes, provenance origin (site name and country), latitude and longitude in decimal degrees, altitude (Alt), the availability of genotypic data and the mean values across all trees of each provenance for the phenotypic traits Budburst, Diameter at breast height (DBH) and Form.

| Site number in Fig. 1 | Code | Provenance name | Country | Long (DD) | Lat (DD) | Alt (m) | Genotypic data | Mean Budburst | Mean DBH | Mean Form |
|-----------------------|------|-----------------|---------|-----------|----------|--------|----------------|--------------|----------|----------|
| 1                     | 185  | Blakeney         | UK      | −2.5      | 51.78    | 76     | No             | 1.08         | 120      | 4.24     |
| 2                     | 217  | Bercé           | France  | 0.39      | 47.81    | 155    | Yes            | 1.59         | 98       | 4.2      |
| 3                     | 237  | Réno Valdieu     | France  | 0.67      | 48.5     | 230    | Yes            | 1.65         | 116      | 4.14     |
| 4                     | 210  | Saint Germain    | France  | 2.08      | 48.9     | 60     | Yes            | 1.71         | 113      | 4.13     |
| 5                     | 194  | Soudrain         | France  | 2.38      | 46.95    | 178    | Yes            | 1.24         | 110      | 4.33     |
| 6                     | 211  | Prémery          | France  | 3.6       | 47.2     | 300    | Yes            | 1.55         | 106      | 4.34     |
| 7                     | 201  | La Haie Renault  | France  | 4.95      | 48.67    | 180    | Yes            | 1.77         | 109      | 4.24     |
| 8                     | 245  | Étangs           | France  | 4.96      | 46.93    | 200    | Yes            | 2.4          | 113      | 4.19     |
| 9                     | 233  | Vachères         | France  | 6.63      | 43.98    | 650    | Yes            | 3.69         | 94       | 4.15     |
| 10                    | 230  | Romersberg       | France  | 6.73      | 48.82    | 220    | No             | 1.06         | 101      | 4.17     |
| 11                    | 250  | Cochern          | Germany | 7.05      | 50.08    | 400    | Yes            | 1.86         | 112      | 4.43     |
| 12                    | 225  | Still            | France  | 7.25      | 48.58    | 688    | Yes            | 1.55         | 105      | 4.56     |
| 13                    | 252  | Johanneskreuz    | Germany | 7.83      | 49.4     | 460    | Yes            | 1.03         | 116      | 4.42     |
| 14                    | 257  | Wolfgang         | Germany | 9.05      | 50.15    | 160    | Yes            | 1.61         | 96       | 3.92     |
| 15                    | 181  | Horbylunde       | Denmark | 9.41      | 56.13    | 80     | Yes            | 0.88         | 89       | 3.99     |
| 16                    | 255  | Spakensehl       | Germany | 10.6      | 52.8     | 115    | Yes            | 0.54         | 102      | 4.74     |
| 17                    | 248  | Kloster-marienberg | Austria | 16.57     | 47.41    | 310    | Yes            | 3.23         | 103      | 3.9      |
| 18                    | 179  | Sycow            | Poland  | 17.93     | 51.18    | 210    | Yes            | 1.42         | 104      | 4.76     |
| 19                    | 249  | Bolu             | Turkey  | 31.67     | 40.92    | 1200   | Yes            | 1.58         | 94       | 4.39     |
| 20                    | 184  | Telavi           | Georgia | 45.47     | 41.88    | 700    | No             | 3.79         | 77       | 3.06     |

Values for the provenance closest to the trial site, Forêt de Bercé, are shown in bold.
the previous year carrying the new leaves). Galls were identified to species and generation in the field using a morphological key (Redfern and Shirley 2002) and our own extensive field experience with this group (Stone et al. 2002; Csóka et al. 2005).

Statistical analysis

Controlling for statistical nonindependence of provenances

One way to address potential statistical nonindependence of populations within species is to incorporate expected
similarities among populations (due to shared common ancestry and/or gene flow) as one or more variance–covariance matrices in the random effect structure of a generalized linear mixed model (Stone et al. 2011). Covariance matrices for nonindependence arising through each of migration and common ancestry can be generated for small numbers of populations using population genomic data (e.g., Hearn et al. 2014). However, for larger numbers of populations (as here), and where more limited genotypic data are available, an alternative is to use a measure of pairwise genetic differentiation between populations such as $F_{ST}$ (Weir and Cockerham 1984) as an estimate of their expected covariance (Stone et al. 2011). Where genetic differentiation between populations correlates strongly with geographic distance, variance–covariance can be estimated as a function of geographic separation between populations.

We used data from 10 putatively neutral microsatellite markers (listed in Data S2; Guichoux et al. 2011) available for 17 of our 20 study provenances (Table 1) to estimate pairwise $F_{ST}$ as implemented in iSTAT version 2.8.3.2 (Weir and Cockerham 1984; Goudet 1995). For comments on the impact of using alternative measures of genetic differentiation related to $F_{ST}$, see Data S2. To allow us to analyse data for all 20 study provenances, we used the highly significant relationship between $F_{ST}$ and geographic distance for the 17 provenances (Mantel test in the adegenet R package, $P = 0.01$ with 999 iterations) to estimate a pseudo-$F_{ST}$ for the full set of 20 provenances, referred to hereafter as $GeoF_{ST}$ (full details are given in Data S2). Variance–covariances between provenances were then estimated as either $1 - F_{ST}$ or $1 - GeoF_{ST}$, such that expected covariance was greatest between provenances with lowest differentiation.

To explore the consequences of using alternative approaches to control for provenance nonindependence, all of the analyses described below were conducted using the following five dataset/covariance matrix combinations: (i) the subset of 17 provenances for which genotypic data were available, using an identity matrix as the provenance covariance matrix. This identity model assumes provenances to be statistically entirely independent (i.e., 0 covariance between all populations); (ii) the same 17 provenances, using a covariance matrix estimated as $1 - F_{ST}$; (iii) the same 17 provenances, using a covariance matrix estimated as $1 - GeoF_{ST}$; (iv) the full set of 20 provenances using an identity model; and (v) the full set of 20 provenances where covariance is estimated as $1 - GeoF_{ST}$. A detailed discussion of this approach to addressing population nonindependence, and values used in the covariance matrices, are provided in the Data S2.

**Statistical modelling of provenance effects on gall abundance**

To investigate the influence of tree provenance on the abundance of each individual gall type, we combined the data for the two survey years and used generalized linear mixed models (GLMMs) to partition the variance of gall counts per shoot into the effects of year (as a fixed effect with two levels) and provenance, soil zone, plot and tree (all as random effects). When applying an identity model for the relationship between provenances, the effect of provenance was represented by a single random effect with a unique independent level for each provenance. In models that incorporated the relationship between provenances, provenance co-variance was represented by two random effects. One incorporated the specified variance–covariance matrix $(1 - F_{ST}$ or $1 - GeoF_{ST})$ to account for provenance nonindependence, while the second incorporated an identity matrix to capture the ‘nugget effect’ (Matheron 1963), in which variation in the response variable exceeds that expected through correlation between levels of the effect. Models were fitted in a Bayesian Markov Chain Monte Carlo (MCMC) framework using the package MCMCglmm version 2.16 (Hadfield 2010) in R version 2.15.1. For fixed effects, prior settings assumed a normal distribution with a mean ($\mu$) of 0 and a variance ($\sigma^2$) of $10^5$. For random effect variances, scaled $F_{-\{1,1\}}$ priors (with scale 1000) were used (Gelman 2006) except for the residual variance for which an inverse-gamma prior was used with shape and scale equal to 0.001. Models were run for 500 000 iterations with a burn-in of 50 000 and parameter sampling every 450 iterations. As the response variable was count data and usually contained many zeros, a Poisson error distribution was applied. Our incorporation of year as a fixed effect allowed for variation in abundance between years, while estimating the relationship between gall abundance and tree phenotypic traits across both years. The proportion of variance explained by the year fixed effect was estimated as the marginal $R^2$ of the model (RsqM), following Nakagawa and Schielzeth (2013) (R code for deriving RsqM from MCMCglmm objects is provided in Data S3). The variance component of each random effect was calculated by dividing its estimated variance parameter by the sum of all the random effect variance parameters, including the residual UNIT term – a random effect with a distinct level for each data unit that is fitted automatically for Poisson distributed models in the MCMCglmm package to account for overdispersion – and then multiplying each by the proportion of variance not explained by the fixed effect (i.e., 1–RsqM). We addressed the effects of multiple tests on false discovery rates – the expected proportion of false discoveries among the rejected hypotheses – following Benjamini and Hochberg (1995) and Benjamini and Yekutieli (2001). The false discovery rate is a less stringent condition than the family-wise error rate, so this approach is more powerful than alternatives such as Bonferroni (Nakagawa 2004).
Effect of tree phenotype

To investigate whether tree phenotypic traits were good predictors of variation in gall abundance, GLMMs were used to model counts of individual gall types with DBH (growth), Form (tree health) or Budburst (spring phenology) as fixed effects. While the influences of tree growth and health on gall abundance are expected to be linear, the local adaptation hypothesis for herbivores predicts that the influence of phenology will be curvilinear, with lower gall abundance on provenances that show budburst either before or after local oaks. The fixed effect Budburst was therefore fitted as both untransformed and squared terms. To account for potential differences in abundance levels between the two study years, a factor year was fitted as a fixed effect.

For each gall type, counts per shoot were initially modelled with year and either Form, DBH, Budburst or both Budburst and Budburst$^2$ as fixed effects. The severity of co-linearity between these variables was assessed by their variance inflation factors (VIFs) using the R package HH version 2.3-23 (Heiberger 2013). To assess the importance of interactions between tree phenotypic traits, a model containing the fixed main effects Form, DBH, Budburst, and their three pairwise interactions was fitted for each gall type. All models were fitted in the MCMCglmm package with the random effect structures and prior specifications described above. The probability that the posterior distribution of the fixed effect coefficients did not include zero (pMCMC) was used to assess the predictive capacity of the fixed effects.

Results

Phenotypic variation among provenances

If provenances are locally adapted, and divergence in selection history between provenance origins and the trial site increases with genetic distance, then we expect more genetically divergent provenances to show higher stress (measured using Form) and lower vigour (measured using DBH). Form shows no linear relationship, but increases in variance in more divergent provenances (Fig. 2A). Only the most genetically divergent provenance, from Telavi in Georgia, showed a markedly lower Form score than other provenances. Plant vigour, measured using DBH, does decrease with increasing genetic divergence (Fig. 2B), consistent with local adaptation. Budburst phenology shows a strong relationship with latitude (Fig. 2C), with more southerly provenances showing earlier budburst: by the time some southern provenance trees have open leaves (scores 4 and 5), northern provenances have yet to burst their buds (scores 0 and 1). We thus expect phenological mismatch with native oaks to correlate with differences in latitude between provenance origins and the trial site.

Herbivore abundance

We recorded approximately 725 000 galls of 20 gall types (seven sexual generation and 13 asexual generation) during the four survey seasons, representing 15 different gallwasp species (Table 2). No acorn galls were recorded. Total counts of individual gall types varied over five orders of magnitude, from just 4 (asexual generation Cynips longiventris) to over 330 000 (asexual generation Neuroterus anthracinus). Gall abundance differed significantly between years for all gall types except Andricus solitarius, with five types exhibiting >10-fold differences (Table 2) and without any consistent pattern between years or seasons. Six of the 20 gall types occurred on <1% of shoots in both survey years and were excluded from further analysis.

Oak provenance effects on gall abundance

The percentage of variance in gall abundance per shoot explained by differences between provenances (expressed as the mean variance component, ranging between 0% and 100%) varied substantially between the statistical models used. In identity models of the 20 provenance dataset (which assume provenances to be fully independent statistically), the percentage of variance in gall abundance explained by provenance ranged from 3% (in sexual generation Neuroterus albipes) to 33% (in asexual generation Neuroterus quercusbaccarum) and was >10% for seven of the 14 gall types (Fig. 3). Because provenances are related, the identity model is unlikely to be appropriate. All models incorporating provenance covariance gave higher estimates of the variance component for provenance effects (see Data S2 for comparison of results across the alternative datasets and models). Incorporation of provenance covariance using the 1−GeoF$_{ST}$ covariance matrix for the same 20 provenance dataset attributes between 38% (sexual generation N. albipes and 95% (asexual generation N. quercusbaccarum) of variance in gall abundance to provenance effects (Fig. 3). The identity models should thus be viewed as providing artificially low estimates of the variation in gall abundance attributable to differences between provenances (Table S2.2, Figure S2.1 in Data S2). On this basis, host tree provenance is an important predictor of the abundance of many gallers, explaining at least 10%, and as much as 95% of the variation in abundance of seven of the 14 gall types tested. Without more knowledge of the form of variance–covariance relationship, it is not possible to give a correct estimate. Importantly for what follows, analysis of fixed effects was not influenced by our choice of variance–covariance model.
Tree phenotypic traits as predictors of variation in gall abundance

Significant relationships between gall abundance per shoot and at least one of the provenance phenotypic traits were identified for all but three of the 14 gall types tested (Table 3). The signs and strengths of relationships were stable across the various datasets and provenance covariance structures (see Data S2 for full comparison of results), and here, we therefore focus on the results of the 1/C0 GeoFST co-variance matrix models for the 20-provenance dataset, identifying where they differ from the identity models for the same dataset (Table 3). The phenotypic variables Form, DBH and Budburst were not confounded, with low VIFs (1.05, 1.02 and 1.02, respectively). Because our analyses yielded only one significant interaction term (Budburst: Form for sexual generation galls of Neuroterus numismalis; mean coefficient = 0.031, pMCMC = 0.014), hereafter we consider each trait in isolation.

Tree health, as measured by Form, significantly predicted the abundance of only one gall type (asexual generation Cynips quercusfolii; Table 3, Fig. 4A) for which healthier trees supported more galls per shoot, as predicted by the plant vigour hypothesis. Tree vigour, as measured by DBH, was a better overall predictor of gall abundance per shoot, with significant relationships for five of the 14 gall types (two sexual and three asexual generations; Table 3, Fig. 4B). For the spring sexual generation galls (Andricus testaceipes and N. numismalis), galls were more abundant on larger trees (Fig. 4B, as predicted by the plant vigour hypothesis). Relationships varied for the three autumn asexual generation gall types: N. albipes were more abundant on larger trees (as predicted by the plant vigour hypothesis), while N. numismalis and N. quercusbaccarum showed the opposite pattern (as predicted by the plant stress hypothesis).

Phenology, as measured by Budburst, was the best overall predictor of gall abundance/shoot, with significant correlations for eight gall types (Table 3, Fig. 4C). Four gall types showed significantly nonlinear relationships, incorporating a significant Budburst squared term. However, no gall type showed peak abundance at an intermediate phenological stage, as expected for the local adaptation hypothesis. Instead, spring sexual generation galls were generally most abundant on provenances showing the earliest budburst (high budburst score), while the summer/autumn asexual

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Table 2. Prevalence, incidence and between-year variation in the abundance of the 20 gall types recorded at La Petite Charnie. Columns show, for each survey year, the total number of each gall type, the mean number of galls per shoot and incidence – the proportion of shoots bearing galls, and the ratio of mean galls per shoot between 2008 and 2009.

| Gall type (species and generation) | 2008 Total galls | Mean galls/shoot | Incidence | 2009 Total galls | Mean galls/shoot | Incidence | Ratio of mean galls/2008:2009 |
|-----------------------------------|-----------------|-----------------|-----------|-----------------|-----------------|-----------|-----------------------------|
| Spring sexual generation surveys  |                 |                 |           |                 |                 |           |                             |
| Andricus inflator                 | 52              | 0.0022          | 0.0018    | 0               | 0               | 0         | –                           |
| Andricus testaceipes              | 973             | 0.041           | 0.0025    | 2291            | 0.095           | 0.067     | 1.23                        |
| Biorhiza pallida                  | 13              | 0.00054         | 0.00046   | 10              | 0.00042         | 0.00042   | 1.31                        |
| Neuroterus albipes                | 822             | 0.034           | 0.03      | 5828            | 0.24            | 0.16      | 1.8                         |
| Neuroterus anthracinus            | 14 278          | 0.59            | 0.35      | 7411            | 0.31            | 0.23      | 1.91                        |
| Neuroterus numismalis             | 1154            | 0.048           | 0.039     | 17 486          | 0.73            | 0.29      | 1.15                        |
| Neuroterus quercusbaccarum        | 5808            | 0.24            | 0.15      | 7834            | 0.33            | 0.2       | 1.14                        |
| Autumn asexual generation surveys |                 |                 |           |                 |                 |           |                             |
| Andricus callidoma                | 4               | 0.00021         | 0.00021   | 45              | 0.0019          | 0.0018    | 1.9                         |
| Andricus fecundatrix              | 334             | 0.017           | 0.013     | 1103            | 0.046           | 0.029     | 1.25                        |
| Andricus glandulai                | 83              | 0.0043          | 0.0037    | 335             | 0.014           | 0.012     | 1.23                        |
| Andricus inflator                 | 44              | 0.0023          | 0.002     | 3               | 0.00013         | 0.00013   | 19.1                        |
| Andricus kollari                  | 2               | 0.0001          | 0.0001    | 43              | 0.0018          | 0.0012    | 1.18                        |
| Andricus solitarius               | 528             | 0.028           | 0.026     | 726             | 0.03            | 0.028     | 1.1                         |
| Cynips divisa                     | 1901            | 0.099           | 0.057     | 184             | 0.0077          | 0.0057    | 13.1                        |
| Cynips longiventris               | 1               | 0.000052        | 0.000052  | 3               | 0.00013         | 0.00013   | 1.25                        |
| Cynips quercusfolii               | 197             | 0.01            | 0.0085    | 456             | 0.019           | 0.013     | 1.2                         |
| N. albipes                        | 45 771          | 2.38            | 0.51      | 27 531          | 1.15            | 0.32      | 2.11                        |
| N. anthracinus                    | 313 507         | 16.33           | 0.99      | 23 820          | 0.99            | 0.29      | 16.1                        |
| N. numismalis                     | 12 194          | 0.64            | 0.037     | 12 124          | 0.51            | 0.028     | 1.21                        |
| N. quercusbaccarum                | 123 708         | 6.44            | 0.56      | 94 830          | 3.95            | 0.49      | 1.61                        |
Influences of oak provenance on the gallwasp guild, and incorporation of provenance covariance

Our results show that oak provenances vary significantly in phenotypic traits likely to influence associated herbivores.
Figure 4 Plots of significant relationships between gall abundance and (A) Form, (B) Diameter at breast height (DBH) and (C) Budburst. Higher Form scores indicate healthier trees, while higher Budburst scores indicate earlier budburst. Distributions of $y$ at intervals across the range of a phenotypic predictor ($x$) were derived from the 1000 stored Markov Chain Monte Carlo (MCMC) samples as: $y_i = \exp(\text{intercept}_i + x_i \text{mean coefficient}) + 0.5x \sum$ random effect variances. Plots show the means (bold lines) and 95% confidence intervals from these distributions for each of the two study years. Dotted vertical lines indicate the mean value of the phenotypic trait for the most local tree provenance (Forêt de Bercé). Plots for models of the asexual generation galls of Neuroterus numismalis could not be produced in this way due to high random effect variances.
and that this in turn is associated with significant variation in the population density (abundance per oak shoot) of multiple gallwasp species. We show that the form of these relationships is robust across models that assume very different patterns of variance–covariance among oak provenances, from identity models assuming total independence to models that estimate covariance based on population genetic and provenance origin data. Our results indicate that models that treat provenances as statistically independent entities, which we know not to be true, consistently underestimate provenance effects on gall densities relative to models in which covariance is specified based on a measure of genetic differentiation ($F_{ST}$ or $GeoF_{ST}$). If these models accurately describe provenance co-variation, then the influence of tree provenance on gall abundance is substantially greater (38–95%) than would be inferred by assuming provenances to be statistically independent. While pairwise $F_{ST}$ and related measures are consistently significantly larger than 0 (Tables S2.1 and S2.3 in Data S2), their absolute values are small, providing little resolution of the relationship of genetic differentiation and geographic distance (see Data S2 for a comprehensive discussion). It is therefore difficult to estimate a ‘true’ value for the impact of provenance and its explanatory power given the nonindependence of populations (Felsenstein 2002; Stone et al. 2011).

The strong provenance effects we show are consistent with the growing body of community genetics work linking host plant genotypic variation to associated trophic levels (Wimp et al. 2005; Bangert and Whitham 2007). A key challenge in this field is to understand the importance of host genetic effects for herbivore communities relative to other ecological factors (Hersch-Green et al. 2011). Working near the northern limit of pedunculate oak ($Quercus robur$) in Finland, Tack et al. (2010) found spatial effects to be more important than tree genotypic effects in structuring a community of herbivorous insects that included several gallwasps. In reciprocal transplant experiments at local and regional scales, Tack et al. (2010) found no effect of tree genotype for any of the three gall types in their study (asexual and sexual generations of $N. quercusbaccarum$ are shared with our study). The contrast between this and the strong host plant provenance effects present at the Petite Charnie trial is likely to reflect contrasting levels of adaptive genetic and phenotypic diversity of the oaks in each study. Although Finnish oak populations show substantial intraspecific genetic diversity (Vakkari et al. 2006), they show substantially less genetic diversity (and we expect also phenotypic diversity) than the much wider geographic array of provenances present at La Petite Charnie (Petit et al. 2002). Our provenance trial study involves a much greater array of genetic diversity than would be expected from a single natural population, a feature shared with the studies of genetically diverse hybrid populations (e.g., Dungey et al. 2000) that often show strong plant genotype effects on associated herbivores (Tack et al. 2010). Generally, the importance of host genetic effects relative to other factors is likely to correlate positively with the adaptive genetic diversity within a tree population present at a site, and we would expect this to be increased in situations where climate matching or other forms of assisted migration are implemented.

**Impacts of between-provenance variation in phenology**

The strongest predictor of between-provenance variation in gallwasp density is variation in spring budburst phenology. Phenological changes have long been recognized as a major component of species responses to climate change that can cascade through associated trophic levels. This has been shown, for example, in work on the impact of changes in oak budburst phenology on abundance of associated leaf-feeding moths and breeding responses of insectivorous birds – particularly $Parus major$ (Van Asch and Visser 2007; Visser 2008; Van Asch et al. 2012; Vatka et al. 2014). A key question in research on these systems has been the extent to which responses of herbivores, and their natural enemies are the result of phenotypic plasticity and local adaptation.

It is striking that although adaptation to individual trees has been demonstrated in oak–gallwasp interactions (Egan and Ott 2007), none of our gall types showed declining abundance with increasing phenological mismatch between native oaks and planted provenances, rejecting our local adaptation hypothesis. We also found the responses of the spring and autumn generations in the gallwasp lifecycle can be very different. Spring galls were more abundant on early budbursting trees, while autumn generations of the same species ($N. numismalis$ and $N. quercusbaccarum$) were more abundant on late budbursting trees. Similar relationships to those we observed were recorded for $Neuroterus$ gallwasp in response to local phenological variation in native UK $Q. robur$ populations by Askew (1962, 1984) and Crawley and Akhteruzzaman (1988); our results show that the same trends can be extended to the wider phenological variation present in a provenance trial. We hypothesize that this contrast between the spring and autumn generations may exist because later budbursting trees may have higher abundance of phenologically young tissues suitable for gall induction. Annual gallwasp phenology may thus be the result of contrasting selection on the two lifecycle generations, perhaps limiting local phenological adaptation in either generation (Singer and Parmesan 2010). The dynamics of such complex lifecycles are poorly understood, and complex bottom-up impacts may apply to other groups with multivoltine or more complex lifecycles, par-
Impacts of between-provenance variation in tree vigour

Tree vigour also explained significant between-provenance variation in oak gall density. In contrast, tree shape – commonly used as an indicator of physiological stress – had little impact. As with phenology, the sign of correlations between gall density and tree vigour varied among gall generations and taxa. Significant correlations with tree vigour were positive for spring sexual generation gall types, while two of three were negative for asexual generation gall types. Support was thus evenly split between the plant vigour plant stress hypothesis (Price 1991; Cornelissen et al. 2008) and the plant stress hypothesis (White 1969, 1974, 2009). One possibility is that the apparent contrast in patterns for spring and autumn generation galls could be driven by the long-known seasonal shift in oak resource investment from growth in the spring to metabolic defence in the summer (Feeny 1970). Highly vigorous trees could then be both best for spring generation galls and worst (=best defended) for autumn generation galls. Mechanisms of oak defence against gallwasps remain poorly understood.

Predicting the impact of climate-matched planting on associated herbivores

Climate matching necessarily requires introduction of phenotypically divergent tree populations. In Europe, climate models predict warmer summers and winters (Giorgi and Coppola 2009; Jenkins et al. 2009), and climate matching will thus always involve introducing more southerly provenances that show earlier budburst (Ducousso et al. 1996; Broadmeadow et al. 2005). Under low and high emissions scenarios for 2050, Broadmeadow et al. (2005) identified climate-matched sites for five UK locations that were 0–2° and 5–12° further south, respectively (Broadmeadow et al. 2005). Our results (Fig. 2C) suggest that while the small latitudinal shifts required under a low emission scenario would probably have little impact on oak phenology, the shifts of >10° of latitude required under a high emission scenario could lead to major phenological differences between native and climate-matched trees, and significantly dissimilar gallwasp assemblages.

The widespread existence of latitude-dependent photoperiod cues in plant phenology (Aitken et al. 2008) suggests that the phenological contrasts between climate-matched and native oak populations will be shared by other plant species. The extent to which any such differences influence associated herbivores is likely to vary according to specific plant–herbivore interactions, reflecting available herbivore phenotypic plasticity and adaptive genetic diversity. We hypothesized that impacts of phenological mismatch will be greatest for interactions in which herbivore fitness is strongly correlated with successful exploitation of a narrow phenological ‘window of opportunity’ (Durant et al. 2007; Van Asch and Visser 2007; Forkner et al. 2008; Netherer and Schopf 2010). This includes many insects that feed on newly flushed leaves (Van Asch and Visser 2007; Forkner et al. 2008). However, we found no evidence of such a window for gallwasps: while they prefer to oviposit (or survive better on) young tissues, they show no strong preference for host trees at a single phenological stage. As discussed above, impacts on herbivores with multivoltine or multihost lifecycles may be complex.

The relationships we found between host plant vigour and herbivore abundance are relevant to two stages in a climate matching strategy: the expected poor performance under current climates of introduced provenances matched to predicted future conditions, and the expected decline in the performance of native provenances as climates move away from any locally adapted range. The slope of the relationship between environmental variables and vigour for a specific provenance will dictate not only the impact on associated trophic levels, but also whether climate matching with that provenance is feasible at all. The challenge for forest managers is to identify a climate-matched provenance that is able to survive as a sapling and young tree before the conditions justifying its selection have been reached and become established before native provenances fail. Climate matching is a recently developed strategy and its success at maintaining forest productivity is untested (Broadmeadow et al. 2005). Current guidelines in UK forestry advocate mixed planting of both matched and local provenances, which increases local resource heterogeneity for herbivores (Hubert and Cottrell 2007), and creates conditions that accelerate both environmental selection on host trees (Lefevre et al. 2014) and host plant-mediated selection on associated organisms.

Assisted migration is increasingly seen as a viable strategy for commercial and conservation settings (Broadmeadow et al. 2005; Stone 2010; Dawson et al. 2011). Issues cited as slowing down the implementation of climate matching as an adaptive forestry strategy include lack of operating procedures, uncertainties over which climate models to use in matching algorithms and risks associated with moving plants outside their current ranges (Williams and Dumroese 2013). Theoretical studies nevertheless suggest that such adaptive-, rather than resilience-targeted strategies may be the best way to safeguard forest systems (Buma and Wessman 2013; Lefevre et al. 2014). We would add that the community impacts of assisted migration need to be carefully considered, particularly so for foundation species such as forest trees. This requires better understanding of the
relative contributions of host plant genotype and other ecological factors to the tree phenotypes exposed to herbivores (Tack et al. 2010). Geographical patterns of local adaptation and tree phenotypic variation have only been mapped in a tiny minority of species (Savolainen et al. 2007; Robakowski et al. 2012; Sork et al. 2013), and studies that target multiple provenance trial locations are required to separate genetic and environmental contributions to provenance phenotypes (Tack et al. 2010). While our study was based on a single trial, the oak phenotypic traits we analysed have been shown to be highly heritable in the sampled provenances (Sinclair 2012), so we expect the provenance effects we found to contain a significant genetic component. Given strong environmental effects on plant traits including phenology, repetition of the work in an array of trials spanning different environments is clearly desirable. Understanding herbivore numerical and adaptive responses then in turn requires understanding the impacts of both plant phenotypes and other ecological processes, including the mortality inflicted on herbivores by their natural enemies (Bidart-Bouzat and Imeh-Nathaniel 2008; Thomson et al. 2010; Van Asch et al. 2012; Rasmann et al. 2014). Much more needs to be done before forest managers can take informed decisions concerning likely impacts of climate matching on forest-associated biodiversity.

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Data archiving statement

Data for this study are available for download from the NERC Environmental Information Data Centre: http://doi.org/10.647.

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Supporting Information

Additional Supporting Information may be found in the online version of this article:

Data S1. Petite Charnie provenance trials.

Data S2. Investigating the issue of non-independence of provenances.
Data S3. Calculating marginal $R^2$ of a MCMCglmm model.
Data S4. Consistency of spring bud-burst phenology between years.
Data S5. An exemplar .csv data frame for use with Data S3.