Six Methods for Latent Moderation Analysis in Marketing Research: A Comparison and Guidelines

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Abstract
It is common in moderation analysis that at least one of the target moderation variables is latent and measured with measurement error. This article compares six methods for latent moderation analysis: multigroup, means, corrected means, factor scores, product indicators, and latent product. It reviews their use in marketing research, describes their assumptions, and compares their performance with Monte Carlo simulations. Several recommendations follow from the results. First, although the means method is the most frequently used method in the review (95% of articles), it should only be used when reliabilities of the moderation variables are close to 1, which is rare. Then, all methods except the multigroup method perform similarly well. Second, the results support using the factor scores method and latent product method when reliabilities are smaller than 1. These methods perform best with parameter and standard error bias less than or equal to 5% under most investigated conditions. Third, specific settings can warrant using the multigroup method (if the moderator is discrete), the corrected means method (if moderation variables are single indicators), and the product indicators method (if indicators are nonnormally distributed). Practical guidelines and sample code for four statistical platforms (SPSS, Stata, R, and Mplus) are provided.

Keywords
moderation analysis, measurement error, research methods

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Investigating the boundary conditions of a phenomenon is central to academic research and crucial for decision makers. In marketing, it commonly involves a latent moderation analysis in which at least one of the target moderation variables is latent and is measured by one or more reflective indicators. Recent examples include Auh et al. (2019), who showed that customer orientation dampens the effect of customer participation on satisfaction (all three variables are latent). Another example is a study by Atasoy and Morewedge (2017) that found greater differences in (latent) perceptions of psychological ownership between physical and digital books (manipulated) when consumers had a stronger need for control (latent trait).

This article focuses on latent moderation analysis and compares six main methods that differ in their approach and assumptions: the multigroup, means, corrected means, factor scores, product indicators, and latent product methods. Table 1 summarizes a literature review of the 1,381 articles published in Journal of Marketing Research, Journal of Marketing, Journal of Consumer Research, and Marketing Science between 2015 and 2020. It shows that the moderation methods have not been equally popular. Among 885 estimated moderation effects in 213 articles, 95% of articles used the means method.

The means method takes unit-weighted mean scores of the indicators without accounting for the remaining measurement error in the scores. Measurement error is the difference between observed and true values of a score (Wooldridge 2015, p. 288). Its magnitude is determined as 1 minus the score’s reliability, which is the proportion of systematic variance in the score with respect to its total variance (Bollen 1989, p. 156). It is known that not accounting for measurement error can severely bias estimates and/or standard errors (Bollen 1989; Cohen et al. 2003; Grewal, Cote, and Baumgartner 2004; Spearman 1904; Wooldridge 2015). Bias is the difference between estimated and true values of a parameter or its standard error (Wooldridge 2015). Thus, the popularity of the means
method is in stark contrast with its reported poor statistical properties in the face of measurement error.

Nevertheless, multiple reasons can explain the common use of the means method. First, reliabilities of measures in the literature are high (a mean of .87 in Table 1). However, as Grewal, Cote, and Baumgartner (2004, p. 528) conclude, “Even when reliability is fairly high by conventional standards, measurement error can be damaging.” One may also overlook that measurement error becomes more severe in latent moderation settings because the reliability of an interaction term is usually lower than the reliability of its components (Busemeyer and Jones 1983; McClelland et al. 2017). Second, researchers might believe that ignoring measurement error leads to underestimated moderation effects and that the means method would therefore be a conservative estimator. However, this is only the case for regressions with a single predictor, in which not accounting for measurement error would bias parameter estimates to zero (Bollen 1989; Cohen et al. 2003; Grewal, Cote, and Baumgartner 2004; Spearman 1904; Wooldridge 2015). The direction and magnitude of bias in models with multiple predictors, even if some are with and some are without measurement error, are more difficult to predict (Bollen 1989; Cohen et al. 2003; Wooldridge 2015). Third, the lack of a comprehensive performance assessment of the six main latent moderation methods hinders an informed use of these methods. This last point motivated this research.

Our objective is to compare the six methods for latent moderation analysis, both theoretically and empirically, and to provide recommendations for their use. First, we describe the six methods and their differences. Second, we use eight Monte Carlo simulation studies to investigate the statistical properties of the methods under a variety of conditions and in terms of four performance criteria (parameter bias, standard error bias, root mean square error [RMSE], and power). The simulations manipulate, respectively, reliability of the measures (Study 1), scale of the indicators (Studies 2a and 2b), correlation between the latent moderation variables (Study 2c), factor loadings (Study 3), and indicator distributions (Study 4a). They show that some methods, specifically the factor scores method and the latent product method, outperform the others. In addition, the simulations examine the effects of misspecification, specifically correlated measurement errors (Study 4b) and ignoring U-shaped (polynomial) effects of the latent variables (Study 4c), and all methods perform worse there. Third, we provide recommendations for future use of the methods and make sample code available for four statistical platforms (SPSS, Stata, R, and Mplus) to implement the methods.

This article makes several recommendations for latent moderation analysis. First, when the reliabilities of the moderation variables are close to 1, five of the six methods perform well; thus the choice of method is at the researcher’s discretion. The corrected means, factor scores, product indicators, and latent product methods have parameter bias under 2% and standard error bias under 5% when the reliability of Y, X, and Z is high at .95 (Study 1). Under these conditions, the parameter bias of the means method is a slightly higher 8% (and standard error bias is 3%), less than the 10% that is considered acceptable (Feingold 2019; Muthén and Muthén 2002). In contrast, the multigroup method has an estimated bias higher than 20% and should be avoided when moderators have continuous indicators.

Second, our results support the use of the factor scores method and the latent product method in situations where reliabilities of the moderation variables are lower than 1. Both methods perform equally well under most investigated conditions, with bias levels lower than 5%. This is the case when reliabilities are between .75 and .95 (Study 1); for seven-, five-, and three-point categorical indicators (Study 2a); when correlations of X with Z range between 0 and .60 (Study 2c); and when indicator loadings are unequal (Study 3). Researchers might base their choice of either method on the availability in their preferred statistical software.

Third, we identify specific settings for which the multigroup method and the product indicators method can be reserved. The multigroup method can be used for a discrete moderator, although the corrected means, factor scores, product indicators, and latent product methods also perform well with biases under 5% (Study 2b). The product indicators method might be chosen over the other methods for nonnormally distributed indicators (parameter bias of 5% if skewness of the moderation variables is 3 and excess kurtosis is 10, at a sample size of 200). Yet, its standard error bias can harm statistical conclusion validity (Study 4a).

| Table 1. Summary of Literature Review. |
|---------------------------------------|
| **Characteristic** | **Finding** |
| Number of articles | 213 |
| Number of studies | 399 |
| Number of moderation effects | 885 |
| Median (SD) sample size across studies | 215 (330,267) |
| Number (%) of articles with: | |
| Multigroup method | 4 (2%) |
| Means method | 203 (95%) |
| Corrected means method | 1 (<1%) |
| Factor scores method | 8 (4%) |
| Product indicators method | 1 (<1%) |
| Latent product method | 1 (<1%) |
| Mean or mode (SD) of data features: | |
| Size of the moderation effect | .16 (.12) |
| Size of the main effects | .18 (.15) |
| Correlation of X with Z | .16 (.16) |
| Reliability of Y, X, and Z | .87 (.09) |
| Number of indicators of Y, X, and Z | 3 (9.97) |
| Number of scale points of y, x, and z | 7 (19.83) |

Notes: Literature review of moderation analyses in the 2015–2020 volumes of Journal of Marketing Research, Journal of Marketing, Journal of Consumer Research, and Marketing Science. Percentages do not sum to 100% because of rounding and use of multiple methods within an article. Effect sizes are correlations. Effect sizes and correlations are reported as mean and standard deviation (SD), and number of indicators and scale points have modes and SD. Dropping one study with extreme sample sizes ([Z-score] > 3.33) results in a median sample size of 215 (SD = 55,647). For details, see Web Appendix A.
Web Appendix B provides an overview of sample code to implement the methods in SPSS, Stata, R, and Mplus, available at an OSF repository (https://osf.io/py7jx/).

Latent Moderation Analysis

Moderation Framework

Assume the following structural latent moderation model:

\[ Y = \beta_1 \times X + \beta_2 \times Z + \beta_3 \times XZ + \zeta, \]

where \( Y \) is the outcome variable, \( X \) is an input variable, \( Z \) is a moderator, and \( \zeta \sim N(0, \sigma^2_\zeta) \) is the structural error term. The parameter \( \beta_3 \) captures the moderation effect, and \( \beta_1 \) and \( \beta_2 \) are main effect parameters of \( X \) and \( Z \), respectively. This article focuses on latent (unobserved) \( Y, X, \) and \( Z \) but also considers the situation where \( Z \) is manifest (observed). We do not consider cases where \( Y, X, \) and \( Z \) are all manifest, as standard methods for moderation analysis can be used in such cases (Cohen et al. 2003; Wooldridge 2015). Without loss of generality, we assume a zero intercept of \( Y \).

The parameters of the latent moderation model cannot be estimated directly because the true scores of \( Y, X, \) and \( Z \) are latent and are reflected in one or more indicator variables that contain measurement error. For exposition, this article focuses on three indicators per latent variable, the mode in the literature review (Table 1). We consider both continuous and ordered categorical indicators (e.g., Likert scales). The measurement model for \( X \) (and analogous for \( Z \) and \( Y \)) is

\[ x = \Lambda_x \times X + \epsilon_x, \]

where \( \Lambda_x \) is a vector of loadings or weights and \( \epsilon_x \sim N(0, \theta_\epsilon) \) refers to the indicator measurement errors with covariance matrix \( \Theta \). In terms of notation, we use lowercase (e.g., \( x \)) for indicators and uppercase (e.g., \( X \)) for latent variables or their approximations with mean or sum scores of indicators (e.g., \( \hat{X} \)) or factor scores (e.g., \( \hat{X} \)).

Definitions of Key Concepts and Method Performance Criteria

This article focuses on three key concepts: latent moderation analysis, measurement error, and reliability, defined in Table 2, Panel A. In addition, Table 2, Panel B, defines four focal performance criteria to compare the methods for latent moderation analysis: parameter bias, standard error bias, RMSE, and power/Type I error. Each reflects a statistical property of the estimators that might be affected by measurement error and might vary across methods. This article mainly focuses on the performance criteria with respect to the moderation effect because it leads in determining the presence of moderation, but we also consider the main effects because the moderation type (i.e., crossing or not) depends on the sign, size, and significance of the moderation and main effects parameters (Cohen et al. 2003).

Parameter bias. Measurement error can bias moderation and main effects. Unbiased estimates are crucial measures of scientific knowledge and might inform the managerial relevance of effects (Eisenid 2015). If \( Y, X, \) and \( Z \) are manifest (and \( X \) and \( Z \) are normally distributed and uncorrelated), the true moderation effect (Cohen et al. 2003) is

\[ \hat{\beta}_3 = \frac{\text{COV}[Y, XZ]}{\text{VAR}[XZ]}, \]

and analogous for the main effects, where COV refers to a covariance and VAR to a variance. However, suppose that \( XZ \) is a product of scores (e.g., means) of the indicators of \( X \) and \( Z \):

\[ XZ = X + \epsilon_{XZ}, \]

where \( \epsilon_{XZ} \) is the measurement error of the product of \( X \) and \( Z \) plus normally distributed and random (independent from all true scores and all other \( \epsilon \) values) measurement error \( \epsilon_{XZ} \). Then \( \text{COV}[Y, XZ] = \text{COV}[Y, X] \) but \( \text{VAR}[XZ] \) is inflated such that the estimated moderation effect \( \hat{\beta}_3 \) (Bollen 1989, pp. 154–59) is

\[ \hat{\beta}_3 = \frac{\text{VAR}[XZ]}{\text{VAR}[XZ] + \text{VAR}[\epsilon_{XZ}]} \times \beta_3 = \rho_{\hat{X}Z} \times \beta_3, \]

where \( \rho_{\hat{X}Z} \) is the reliability of \( \hat{X}Z \), or in other words, the proportion of systematic variance in \( \hat{X}Z \). Thus, unless \( \hat{X}Z \) is free of measurement error (i.e., \( \rho_{\hat{X}Z} = 1 \)), the estimated moderation effect is biased toward zero, and the magnitude depends on the reliability of the product. These results are analogous for the main effects if \( X \) and \( Z \) are uncorrelated, but the direction and the magnitude of bias for all parameters becomes more difficult to determine for correlated predictors. Moreover, bias due to measurement error in variables might carry over to parameter estimates of other variables in the model, even if they do not contain measurement error. Yet, measurement error in \( Y \) does not bias moderation effects but might attenuate \( R^2 \) (Bollen 1989; Cohen et al. 2003; Wooldridge 2015).

Bias due to measurement error is not specific to latent moderation analysis. Yet it can be more severe in this setting because product terms typically have a lower reliability than their components.1 The reliability of a product of \( X \) and \( \hat{Z} \) (Busemeyer and Jones 1983, Equation 10) is

\[ \rho_{XZ} = \rho_X \times \rho_Z + \rho_{XZ}^2 \frac{1}{1 + \rho_{XZ}^2}, \]

where \( \rho_{XZ}^2 \) is the squared correlation between the scores of \( X \) and \( Z \). For example, if \( X \) and \( \hat{Z} \) have a reliability of .85 and are correlated at .20, the reliability of the product term is a

1 A reliability estimator of \( X^2 \) is the square of the reliability of \( X \) (Dimitruk et al. 2007), so by definition it is lower than the reliability of \( X \) and usually lower than the reliability of \( X \) and \( Z \) (if the reliability of \( Z \) equals the reliability of \( X \) unless \( X \) and \( Z \) are uncorrelated).
much lower .73. However, a higher correlation between $\bar{X}$ and $\bar{Z}$ increases $\rho_{XZ}$ and increases the power of the estimated moderation effect (McClelland et al. 2017).

**Standard error bias.** Measurement error can also bias standard errors (Bollen 1989; Cohen et al. 2003; Van Smeden, Lash, and Groenwold 2019; Wooldridge 2015). Unbiased standard errors are crucial for valid moderation tests and a valid assessment of the uncertainty of moderation estimates more generally. Note that correcting for measurement error increases standard errors, even if they are unbiased. For correlations, a reasonable approximation for the standard error increase due to the correction is the magnitude by which the correlation is biased downward due to measurement error (Hunter and Schmidt 2004, p. 96). However, standard errors are complex functions of the model, the size of the effect, the sample size, the measure reliabilities, and correlations among predictors (Charles 2005; Yuan, Cheng, and Zhang 2010).

**Root mean square error.** The RMSE is based on the sum of the squared bias and the variance of a parameter. It summarizes parameter recovery (lower is better). It can also be used to choose between unbiased estimators. The method with the lowest RMSE (i.e., lowest parameter uncertainty) among unbiased estimators is preferred. Accounting for measurement error decreases parameter bias and thus decreases RMSE. At the same time, the measurement error correction might increase the RMSE due to the larger standard error. The net effect on RMSE is difficult to predict.

**Power and Type I error.** Power and Type I error are the probability that a parameter of interest is found to be statistically significant (Cohen 1988, p. 1). High power is crucial to find effects if they truly are nonzero. Measurement error decreases power and thus increases required sample sizes (Grewal, Cote, and Baumgartner 2004). If the true parameter is zero, the analogue to power is Type I error. Minimizing it prevents false positive
Figure 1. Method visualizations and model equations.

Notes: The "steps" denote whether the measurement and structural models are estimated separately (in two steps) or not. The visualizations have three indicators for Y, X, and Z for exposition. Circles are latent variables, and boxes are manifest indicators. Unidirectional arrows refer to loadings λ and regression paths β. Then ζ's are structural error terms, omitted from visualizations for exposition, and ε's are measurement errors. Error variances, latent variances, and covariances between explanatory variables X, Z, and XZ are omitted for brevity. Superscript g refers to a discrete grouping variable, and the triangle "1" is an intercept α (Panel A), bars (e.g., X) denote means (Panels B and C), and hats (e.g., X) denote estimated factor scores (Panel D). Panel E uses the "matched pairs" strategy to form three product indicators but readily extends to other indicator pairings. In Panel F, the dot connecting X and Z refers to the moderation effect being inferred from the joint distribution of the indicators of X and Z and not based on observed product terms of X and Z and/or their indicators (Muthén and Muthén 2019).

(continued)
results. Power and RMSE complement each other. For instance, a high upward parameter bias can lead to a high power, but RMSE would detect that the estimator is problematic. Among unbiased methods, both RMSE and power should provide qualitatively similar results.

**Six Methods for Latent Moderation Analysis**

Figure 1 visualizes the six methods for latent moderation analysis and provides model equations. Table 3 summarizes the assumptions of the methods.

**Method 1: Multigroup**

This method estimates separate models for discrete subgroups based on the moderator. We focus on two groups for exposition and because the use of two groups is common in moderation analyses (37% of moderation variables in the literature review). The structural model for each group \( g \) is

\[
Y_g = \alpha_g + \beta_{g1} \times X_g + \beta_{g2} \times Z + \beta_{g3} \times XZ + \zeta_g.
\]

It does not include an interaction term but estimates a \( \beta_{g1} \) parameter for each group. The main effect of \( Z \) is derived from the
intercept $\alpha$. Constraining $\beta_1$ to be equal across groups and testing that model against one with a group-specific $\beta_1$ tests moderation. Measurement models as in Equation 2 can be specified for $Y$ and $X$. Grouping is straightforward for a discrete $Z$, such as different countries or different experimental manipulations. Yet when $Z$ is continuous, grouping requires discretization based on a median or other split. Such discretization uses partial information in $Z$ and adds measurement error to the grouping variable (Irwin and McClelland 2001, 2003).

Method 2: Means

This method uses unit-weighted mean (or sum) scores of the indicators. Although mean scores can be used without estimating a measurement model, McNeish and Wolf (2020) show that unit-weighted means are analogous to assuming a parallel measurement model that constrains indicators to be equally weighted with equal measurement error variances. The structural model then uses the mean scores to estimate the moderation effect without accounting for measurement error in the mean scores:

$$\tilde{Y} = \beta_1 \times \bar{X} + \beta_2 \times \bar{Z} + \beta_3 \times \bar{X}\bar{Z} + \zeta. \quad (8)$$

The means $\bar{X}$ and $\bar{Z}$ can be mean-centered prior to computing the interaction term $\bar{X}\bar{Z}$ to facilitate interpretation and reduce unessential multicollinearity (Cohen et al. 2003; Irwin and McClelland 2001).

Method 3: Corrected Means

This method uses a product of mean scores, as the means method does, but corrects for measurement error in the scores by using reliability estimates. A measurement model as in Equation 2 can be used but with loadings and measurement errors fixed for identification (Bollen 1989). For example, for $XZ$, the loading is $\lambda_{XZ} = 1$ and the error variance is $\sigma_{e_{XZ}}^2 = (1 - \rho_{XZ}) \times \sigma_{XZ}^2$, where $\sigma_{XZ}^2$ is the variance of $XZ$, and $\rho_{XZ}$ is its reliability. Reliabilities of $Y$, $X$, and $Z$ can be estimated with estimators such as Cronbach’s alpha, assuming

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**Table 3. Overview of Method Assumptions.**

| Assumption                                      | Multigroup Method | Means Method | Corrected Means Method | Factor Scores Method | Product Indicators Method | Latent Product Method |
|------------------------------------------------|-------------------|--------------|------------------------|----------------------|--------------------------|----------------------|
| **Measurement Model**                           |                   |              |                        |                      |                          |                      |
| Indicator distribution                          |                   |              |                        |                      |                          |                      |
| $x \sim \text{MVN}(\mu_x, \Sigma_x)$           | Yes               | —            | Yes                    | Yes                  | Yes                      | Yes                  |
| $z \sim \text{MVN}(\mu_z, \Sigma_z)$           | No, discrete      | —            | Yes                    | Yes                  | Yes                      | Yes                  |
| $xz \sim \text{MVN}(\mu_{xz}, \Sigma_{xz})$    | —                 | —            | —                      | —                    | —                        | —                    |
| Account for implied nonnormality in $y$         | No                | No           | No                     | No                   | No                       | Yes                  |
| **Indicator Measurement Errors**                |                   |              |                        |                      |                          |                      |
| All $e_x \sim \text{MVN}(0, \Theta_x)$ freely estimated | Yes             | No           | fixed and equal | No, fixed and equal | Yes                      | Yes                  |
| All $e_z \sim \text{MVN}(0, \Theta_z)$ freely estimated | No, fixed and equal | No, fixed and equal | —                      | —                    | —                        | —                    |
| $e_{xz} \sim \text{MVN}(0, \Theta_{xz})$ freely estimated | —                 | —            | —                      | —                    | —                        | —                    |
| All $e_y \sim \text{MVN}(0, \Theta_y)$ freely estimated | Yes             | No           | fixed and equal | No, fixed and equal | Yes                      | Yes                  |
| **Indicator Loadings**                          |                   |              |                        |                      |                          |                      |
| All $\Lambda_x$ freely estimated                | Yes               | No           | fixed and equal | No, fixed and equal | Yes                      | Yes                  |
| All $\Lambda_z$ freely estimated                | No, fixed and equal | No, fixed and equal | —                      | —                    | —                        | —                    |
| All $\Lambda_{xz}$ freely estimated             | —                 | —            | —                      | —                    | —                        | —                    |
| All $\Lambda_y$ freely estimated                | Yes               | No           | fixed and equal | No, fixed and equal | Yes                      | Yes                  |
| **Structural Model**                            |                   |              |                        |                      |                          |                      |
| $\zeta \sim \text{N}(0, \sigma^2_\zeta)$, uncorrelated with $y$, $x$, $z$, $X$, $Z$, and all $\Theta$ | Yes             | Yes          | Yes                    | Yes                  | Yes                      | Yes                  |

*The product indicators method freely estimates the loadings and measurement errors of the product indicators, but using “matched pairs” assumes that all product indicators are equally good representatives of the latent interaction factor $XZ$ because the moderation result might depend on the choice of indicator pairs if indicators are not equally good, which is undesirable (Foldnes and Hagvet 2014; Marsh, Wen, and Hau 2004).

Notes: All methods except the latent product method use standard maximum likelihood estimation, which uses the expectation maximization algorithm that converges to maximum likelihood estimates (Dempster, Laird, and Rubin 1977; Klein and Moosbrugger 2000). $\text{MVN}(\cdot)$ is the multivariate normal distribution and $\text{N}(\cdot)$ is the normal distribution. The “—” denotes that the assumption is not applicable; that is, the means method does not directly use a measurement model so it does not assume a distribution of the indicators. Similarly, the multigroup and latent product methods do not use manifest interactions or product terms to estimate the moderation effect; only the product indicators method uses products of indicators in the measurement model.
unit-weighted indicators. Then $p_{XZ}$ can be estimated with Equation 6. The structural model relates the latent variables as in Equation 1. Statistically, the mean scores across multiple indicators are single indicators of the latent variables. Thus, the corrected means method can also be used for single-indicator measures if their reliability can be estimated (e.g., Pieters 2017, pp. 699–700).

**Method 4: Factor Scores**

This method uses factor scores that estimate the latent variable scores with linear combinations of the indicators. The first step extracts factor scores from measurement models as in Equation 2 that freely estimate measurement errors and loadings. The second step regresses factor scores of Y on those of X, Z, and the product:

$$\hat{Y} = \beta_1 \times \hat{X} + \beta_2 \times \hat{Z} + \beta_3 \times XZ + \zeta.$$  \hspace{1cm} (9)

There are multiple ways to estimate factor scores. In the context of nonmoderation models, Skrondal and Laake (2001) and Devlieger, Mayer, and Rosseel (2016) have shown that using Bartlett factor scores for outcomes and regression factor scores for predictors produces estimates without bias:

$$\hat{F}_{\text{Bartlett}} = D\Theta^{-2}\Lambda(\Lambda'\Theta^{-2}\Lambda)^{-1},$$  \hspace{1cm} (10)

$$\hat{F}_{\text{Regression}} = D\Sigma_{(o)}^{-1}\Lambda\Phi,$$  \hspace{1cm} (11)

where $D$ is a matrix of indicator-level data, $\Theta$ is the variance covariance matrix of the indicator measurement errors, $\Lambda$ is the matrix of estimated loadings, $\Sigma_{(o)}$ is the observed covariance matrix of the indicators, and $\Phi$ is the variance covariance matrix of the latent variables (Lastovicka and Thamodaran 1991). Bartlett factor scores account for measurement error in Y, and regression factor scores account for measurement error in the predictors; combining these factor scores recovers the parameters in nonmoderation models without parameter bias (Devlieger, Mayer, and Rosseel 2016; Skrondal and Laake 2001). We apply this to the context of latent moderation.

There are several ways to specify the measurement models. Measurement models for Y, X, and Z can be estimated jointly or separately with confirmatory factor analysis (CFA) or exploratory factor analysis (EFA) estimated with maximum likelihood. Skrondal and Laake (2001) have shown that separate factor analyses for Y (1-CFA or unrotated 1-EFA) and the predictors are necessary to avoid parameter bias. The predictors need to be combined in a joint confirmatory factor analysis (2-CFA of X and Z) because the 2-CFA accounts for the factor correlation of X with Z. Skrondal and Laake (2001, pp. 572–73) then show with analytical proofs that estimates are unbiased if the correlation is accounted for. Web Appendix C has additional details, including the extension to (moderated) mediation models.

**Method 5: Product Indicators**

Method 5 specifies a measurement model analogous to Equation 2 for products of indicators, while simultaneously estimating the structural model of Equation 1. There are several ways to specify this model. They differ in the product indicators that are paired for moderation analysis and the constraints that are used to estimate the model. Early on, Kenny and Judd (1984) proposed using a measurement model of product indicators that required multiple constraints on the indicator loadings and measurement error variances. Foldnes and Hagtvet (2014) showed with simulation studies and real-world data that there might be considerable variation in moderation estimates depending on the method used to pair indicators. Using a single pair of indicators uses limited information (Jöreskog and Yang 1996), whereas using all pairs of indicators uses all information but might lead to overly complex models (Marsh, Wen, and Hau 2004). Marsh, Wen, and Hau (2004) proposed a compromise “matched pairs” approach, using all indicators of X and Z but each indicator only once. This approach trades off the use of all indicators while limiting model complexity—avoiding correlated measurement errors for pairs that have common components—with acceptable bias and variance implications (Marsh, Wen, and Hau 2004). Lin et al. (2010) show that using matched pairs and double mean-centering the indicator pairs works well. It avoids the need for constraints, other than those for identification, on the indicator loadings and measurement error variances.

A Web of Science citation analysis signals that the product indicators method is rarely used in the focal journals of the literature review, even beyond the included volumes. The three citations of Marsh, Wen, and Hau (2004) apply the method, whereas three of the four citations of Kenny and Judd (1984) refer to its methodological contribution without application. Moreover, the matched pairs approach in Marsh, Wen, and Hau has accumulated more total citations (596) within and outside the marketing domain than other approaches have (at the time of writing: 588 citations of Kenny and Judd [1984], 272 citations of Jöreskog and Yang [1996], 70 citations of Lin et al. [2010], and 13 citations of Foldnes and Hagtvet [2014]).

Accordingly, we use the matched pairs approach with double mean-centering to represent the product indicators method here. In our running example where both X and Z have three indicators, taking matched pairs results in three product indicators of mean-centered variables, for example, $x_1z_1$, $x_2z_2$, and $x_3z_3$ (Marsh, Wen, and Hau 2004), that are subsequently mean-centered once more (Lin et al. 2010).

**Method 6: Latent Product**

This method estimates the moderation effect from the latent product of X and Z (Klein and Moosbrugger 2000). The latent product method is motivated by the nonnormality in Y that is due to the moderation specification (Klein and 2017).
Moosbrugger 2000). Products of variables (e.g., XZ) are usually nonnormally distributed, even if their components (here X and Z) are normally distributed. Because Y is a function of the nonnormally distributed XZ if there is a nonzero moderation effect, it is also nonnormally distributed (Moosbrugger, Schermelleh-Engel, and Klein 1997). Web Appendix D provides further details and an illustrative example. The latent product method takes the nonnormality in Y directly into account. It is therefore based on an analysis of the indicator distribution and uses the raw data for estimation, unlike the other methods for which the observed covariance matrix is sufficient. The nonnormal indicator distribution can be approximated by a weighted sum or finite mixture of normal distributions (Klein and Moosbrugger 2000). The mixture distribution then becomes a tool to estimate the moderation effect from the latent product of the latent X and Z. Web Appendix E provides further details.

Commonalities and Differences Between the Methods

In terms of commonalities between the six methods, they all rely on the same estimation approach. All methods except for the latent product method use standard maximum likelihood estimation (Bollen 1989). The latent product method uses an expectation maximization algorithm that converges to maximum likelihood estimates too (Klein and Moosbrugger 2000), even though expectation maximization can be computationally intensive and more sensitive to local maxima of the likelihood (Dempster, Laird, and Rubin 1977).

The structural moderation models of five of the six methods (all except the multigroup method) are virtually identical. The crucial difference is in the specification and assumptions of the measurement model (Table 3). The means method takes unit-weighted mean scores of the indicators that assume a parallel measurement model (McNeish and Wolf 2020). The means method does not account for the remaining measurement error in the scores. The corrected means method accounts for this shortcoming of the means method by fixing the amount of measurement error in the variables on the basis of reliability estimates. Yet, it maintains the assumptions of a parallel measurement model. The equal indicator weighting biases reliability estimates downward and therefore might lead to upward parameter bias in the moderation effect even if measurement error is accounted for (McNeish and Wolf 2020).

Whereas the means method and corrected means method assume equally weighted indicators, the measurement models of the factor scores method, the product indicators method, and the latent product method freely estimate the loadings and measurement error variances. There are three differences between these methods. First, the factor scores method is a two-step approach that separately estimates measurement and structural models, whereas the product indicators method and the latent product method estimate the measurement model and the moderation effect simultaneously. Second, although the factor scores method and the latent product method use a product of latent variables or their scores in the case of the factor scores method, the product indicators method uses products of matched pairs of indicators. This method assumes that the product indicators are representative of all possible pairs, essentially assuming equally weighted indicators. There is considerable variation in moderation estimates as a result of different indicator pairings if indicators are not equally good, which is undesirable (Foldnes and Hagtvet 2014; Marsh, Wen, and Hau 2004). Third, the latent product method is the only approach that accounts for the nonnormally distributed indicators of Y due to the interaction (Klein and Moosbrugger 2000). However, it maintains the assumption of normally distributed indicators of X and Z, as do the factor scores method and the product indicators method. Yet, interestingly, the product indicators method uses products of indicators that rarely meet the assumption of being normally distributed because products are usually nonnormally distributed even if their components are normally distributed (Moosbrugger, Schermelleh-Engel, and Klein 1997; Oliveira, Oliveira, and Seijas-Macias 2016). Web Appendix D provides further details.

The multigroup method can include measurement models for the indicators of Y and X to account for indicator measurement error but does not rely on a product of variables and estimates models for discrete subgroups based on the moderators. Although naturally discrete moderators—such as different countries, owners of different brands, genders, experimental manipulations, and so on—can readily be used as grouping variables, grouping by discretizing continuous moderators adds measurement error to the grouping variable and can lead to parameter bias and a decrease of power (Irwin and McClelland 2001, 2003).

In summary, the six methods for latent moderation analysis are all based on maximum likelihood estimation. The main differences are in their approach and assumptions of the measurement model.

Overview of Monte Carlo Simulation Studies

We conduct Monte Carlo simulations to compare the statistical properties of the latent moderation methods across conditions. We use simulations because method performance and impact of design factor on method performance are difficult to derive analytically (Muthén and Muthén 2002; Skrondal 2000).

Summary of Studies

Table 4 summarizes the designs of eight Monte Carlo simulation studies (Studies 1, 2a–c, 3, and 4a–c) that focus on a variety of conditions. All studies, unless indicated otherwise, are under the following conditions. They generate standard normally distributed Y, X, and Z. Data generation is based on values from the literature review as much as possible, thus mimicking real-world situations (Table 1). The latent Y, X, and Z
Table 4. Summary of Study Designs.

| Study                  | Methods Compared | Reliability of Y, X, and Z | Indicator Scale of y, x, and z | Correlation of X with Z | Indicator Loadings | Distribution of x and z | Indicator Measurement Errors | Structural Model Specification |
|------------------------|------------------|-----------------------------|--------------------------------|-------------------------|--------------------|--------------------------|-------------------------------|--------------------------------|
| Study 1: Reliability of measures | 1–6              | .95, .85, .75               | Continuous                     | .20                     | Equal              | Normal                   | Uncorrelated                  | Correctly specified            |
| Study 2a: Ordered categorical indicators | 1–6              | .85                          | Ordered categorical (seven-, five-, three-point scales) | .20                     | Equal              | Ordered categorical      | Uncorrelated                  | Correctly specified            |
| Study 2b: Discrete moderator | 1–6              | .95, .85, .75               | y and x: continuous z: discrete (binary) | 0                      | Equal              | x: normal z: discrete (binary) | Uncorrelated                  | Correctly specified            |
| Study 2c: Correlation of X with Z | 1–6              | .85                          | Continuous                      | 0, .20, .40, .60       | Equal              | Normal                   | Uncorrelated                  | Correctly specified            |
| Study 3: Unequal indicator loadings | 3–6              | .85                          | Continuous                      | .20                     | Unequal (1, 1.5, .50) | Normal                   | Uncorrelated                  | Correctly specified            |
| Study 4a: Nonnormally distributed indicators | 3–6              | .85                          | Continuous                      | .20                     | Equal              | Nonnormal\(^a\)           | Uncorrelated                  | Correctly specified            |
| Study 4b: Correlated measurement errors | 3–6              | .85                          | Continuous                      | .20                     | Equal              | Correlated .50 (x with y, x with z, x with x)\(^b\) | Uncorrelated                  | Correctly specified            |
| Study 4c: Misspecification of the structural model | 3–6              | .85                          | Continuous                      | 0, .20, .40, .60       | Equal              | Normal                   | Uncorrelated                  | Misspecified\(^c\)             |

\(^a\)Study 4a has nonnormality in x and z due to nonnormality in X and Z (skewness and excess kurtosis for X and Z: 1 and 2, or 3 and 10).

\(^b\)In Study 4b, the condition “x with x” means that indicators of X are .50 intercorrelated (with other x indicators).

\(^c\)Study 4c generates a polynomial of X (Y = β₁ × X + β₂ × Z + β₄ × X²) and estimates the parameters with Equation 1.

Notes: Boldface denotes the focus of each study. Methods are 1: multigroup, 2: means, 3: corrected means, 4: factor scores, 5: product indicators, and 6: latent product. Studies 1 and 2a–c investigate all six methods. Studies 3 and 4a–c drop the multigroup and means methods (methods 1 and 2) because of their performance in Studies 1 and 2a–2c. All studies vary the sample size: 100, 150, 200, 300, 500, 750, and 1,500.
variables have three indicators (the most common in the literature review) that are equally good. Reliabilities of Y, X, and Z are .85, which is about the mean in our literature review and the mean in a recent review of mediation analyses (Pieters 2017). The moderation and main effect sizes are .20, which are about the mean values in the literature review and are small-to-medium effects (Cohen 1988). The correlation between X and Z is .20, about the mean in the literature review. Sample sizes are 100, 150, 200 (about the median in the literature review), 300, 500, 750, or 1,500. About 82% of the studies in the literature review have sample sizes between 100 and 1,500.

For each study, we generate 5,000 replications (data sets) per cell in R (R Core Team 2020) using common random number seeds to increase precision and for reproducibility (Skronal 2000). The R package lavaan (Rosseel 2012) implements all methods except for the latent product method, for which we use Mplus 8.3 (Muthén and Muthén 2019) from R via MplusAutomation (Hallquist and Wiley 2018). The OSF repository at https://osf.io/py7jx/ has simulation code for all studies.

**Method Performance Criteria**

Table 2, Panel B, has the operationalizations of the performance criteria to compare the methods. We calculate parameter bias by taking the deviations of the estimated main or moderation effect parameter $\hat{\beta}$ from its true value $\beta$ and dividing by the true value such that the bias is on a percentage scale. We then take the mean across Monte Carlo replications. Similarly, standard error bias is the mean deviation of the estimated standard error from the true standard error, of which the standard deviation of the estimated parameter across replications is an estimate (Muthén and Muthén 2002). Then, RMSE is the square root of the sum of the squared parameter bias and estimated variance (squared standard error), and an estimate of power (or Type I error if the true parameter is zero) is the percentage of Monte Carlo replications for which the parameter of interest is statistically significant at two-tailed $p \leq .05$.

We evaluate the methods as follows. We first calculate biases in parameters and standard errors and retain the unbiased methods. Common acceptable levels of absolute parameter bias are $\leq 10\%$ and $\leq 5\%$ for standard error bias (Feingold 2019; Muthén and Muthén 2002). For the methods that meet these criteria, we consider RMSE and power. However, these criteria are not interpretable for biased methods because downward standard error bias can lead to low RMSE and upward parameter bias can lead to high power. Common thresholds are $\geq 80\%$ for power and $\leq 5\%$ Type I error (Cohen 1988; Muthén and Muthén 2002). Panel B of Table 2 summarizes these thresholds.

Table 5 summarizes the performance criteria for all methods across the conditions for each study at about the median sample size of 200 in the literature review. Web Appendices F–M and the material on the OSF repository (https://osf.io/py7jx/) provide detailed results for the moderation and main effects.

### Study 1: Reliability of Measures

#### Design

Study 1 focuses on measure reliability as a determinant of method performance. The design is 6 (method) $\times$ 7 (sample size) $\times$ 3 (reliability of Y, X, and Z: .95, .85, or .75). The reliability levels are approximately the mean in the literature review, plus and minus one standard deviation (Table 1). These levels are respectively excellent, good, and acceptable reliability (Peterson 1994). We expect that the multigroup method is biased and has a high RMSE and low statistical power because discretizing the continuous indicators of the moderator adds measurement error. We expect the means method to be biased, but we expect the bias to decrease when the reliability increases. In contrast, the latent product method should recover parameters well. An open question is whether the corrected means method, the factor scores method, and the product indicators method perform similarly to the latent product method. Moreover, it is unclear how these methods perform at lower measure reliabilities (i.e., .75) and/or in smaller samples (e.g., 100 observations).

#### Results

Panels A–D in Figure 2 plot performance of the moderation effect estimates (y-axis) across sample sizes (x-axis) for each method (symbols) and across measure reliability levels (.75 in the left plot, .85 in the center plot, and .95 in the right plot of each panel). Overall, methods perform better and more similarly to each other when measure reliability and sample size increase. However, there are several key performance differences between methods.

**Parameter bias (Panel A).** The multigroup method is biased, even at high reliability levels of .95 and large sample sizes (e.g., 1,500), with a bias of about 20%. Similarly, the means method is biased by 41% and 26% at reliabilities of .75 and .85, respectively. Increasing sample size does not reduce bias, making the multigroup method and the means method inconsistent estimators (Wooldridge 2015, p. 287). Yet, the bias of the means method at a reliability of .95 is 8%, which can be acceptable (Table 2). At that reliability, differences between methods become smaller. The corrected means, factor scores, product indicators, and latent product methods have biases of about 1%–2%. Differences between methods become larger at lower reliabilities. The product indicators method is unbiased only at larger sample sizes (e.g., ≥300), when reliabilities are .75. Overall, the corrected means method, factor scores method, and latent product method are unbiased (parameter bias below 6% across reliabilities and at a sample size of 200).

**Standard error bias (Panel B).** All methods except the product indicators method have standard error biases under 5% for samples of at least 200 observations. The product indicators method has biased standard errors (up to about 33% at a reliability of .75 and sample size of 100) when measure reliability is
| Study and Condition | Parameter Bias (in %) | SE Bias (in %) | RMSE | Power/Type I Error (in %) |
|---------------------|-----------------------|---------------|------|--------------------------|
| **Study 1: Reliability of Measures** | | | | |
| Reliability of Y, X, and Z is .95 | 37 | 8 | 2 | 1 | 1 | 4 | 3 | 4 | 1 | 5 | 3 | .34 | .29 | .31 | .30 | .31 | .31 | 60 | 80 | 80 | 80 | 80 | 80 |
| Reliability of Y, X, and Z is .85 | 40 | 26 | 3 | 2 | 4 | 2 | 4 | 3 | 4 | 2 | 10 | 3 | .39 | .31 | .37 | .35 | .39 | .37 | 46 | 66 | 65 | 65 | 61 | 64 |
| Reliability of Y, X, and Z is .75 | 43 | 41 | 6 | 3 | 13 | 4 | 5 | 3 | 4 | 2 | 22 | 4 | .45 | .34 | .44 | .41 | 1.09 | .45 | 34 | 52 | 50 | 51 | 37 | 48 |
| **Study 2a: Ordered Categorical Indicators** | | | | |
| Seven-point ordered categorical scales | 34 | 27 | 1 | 3 | 1 | 2 | 4 | 3 | 3 | 1 | 5 | 2 | .40 | .32 | .38 | .36 | .39 | .38 | 42 | 60 | 59 | 60 | 58 | 58 |
| Five-point ordered categorical scales | 54 | 29 | 1 | 4 | 1 | 2 | 6 | 2 | 3 | 1 | 5 | 1 | .41 | .32 | .39 | .37 | .40 | .39 | 40 | 58 | 57 | 58 | 54 | 56 |
| Three-point ordered categorical scales | 73 | 38 | 36 | 5 | 8 | 4 | 13 | 8 | 43 | 2 | 21 | 3 | .46 | .34 | .51 | .40 | .47 | .43 | 33 | 49 | 46 | 48 | 43 | 45 |
| **Study 2b: Discrete Moderator** | | | | |
| Reliability of Y, X, and Z is .95 | 1 | 5 | <1 | 1 | 1 | <1 | 2 | 3 | 4 | 1 | 4 | 1 | .30 | .29 | .30 | .29 | .30 | .30 | 81 | 81 | 81 | 81 | 81 | 81 |
| Reliability of Y, X, and Z is .85 | 1 | 15 | 1 | 2 | 1 | 1 | 3 | 3 | 3 | 1 | 4 | 3 | .35 | .31 | .33 | .33 | .34 | .34 | 70 | 71 | 70 | 71 | 70 | 70 |
| Reliability of Y, X, and Z is .75 | 2 | 25 | 2 | 3 | 2 | 2 | 5 | 3 | 3 | 1 | 5 | 4 | .40 | .33 | .38 | .37 | .39 | .39 | 58 | 60 | 59 | 59 | 58 | 58 |
| **Study 2c: Correlation of X with Z** | | | | |
| Correlation X with Z is 0 | 27 | 27 | 3 | 2 | 4 | 1 | 4 | 3 | 3 | 1 | 9 | 2 | .37 | .31 | .37 | .35 | .39 | .37 | 44 | 64 | 63 | 63 | 59 | 62 |
| Correlation X with Z is .20 | 40 | 26 | 3 | 2 | 4 | 2 | 4 | 3 | 4 | 2 | 10 | 3 | .39 | .31 | .37 | .35 | .39 | .37 | 46 | 66 | 65 | 65 | 61 | 64 |
| Correlation X with Z is .40 | 58 | 24 | 2 | 2 | 4 | 2 | 5 | 3 | 3 | 2 | 10 | 3 | .43 | .31 | .38 | .36 | .40 | .38 | 50 | 72 | 71 | 71 | 68 | 71 |
| Correlation X with Z is .60 | 86 | 21 | 2 | 2 | 3 | 1 | 5 | 4 | 4 | 2 | 2 | 10 | 3 | .50 | .31 | .41 | .40 | .43 | .42 | 57 | 79 | 79 | 79 | 76 | 79 |
| **Study 3: Unequal Indicator Loadings** | | | | |
| Indicator loadings are unequal | — | — | 25 | 3 | 4 | 4 | — | — | 4 | 3 | 12 | 5 | — | — | .43 | .34 | .38 | .37 | — | — | 67 | 69 | 67 | 68 |
| **Study 4a: Nonnormally Distributed Indicators** | | | | |
| x and z are moderately nonnormally distributed (X and Z skewness is 1, excess kurtosis is 2) | — | — | 5 | 3 | 5 | 5 | — | — | 6 | 5 | 16 | 4 | — | — | .38 | .35 | .41 | .38 | — | — | 66 | 69 | 62 | 68 |
| x and z are severely nonnormally distributed (X and Z skewness is 3, excess kurtosis is 10) | — | — | 19 | 14 | 5 | 14 | — | — | 12 | 13 | 32 | 7 | — | — | .40 | .34 | .43 | .37 | — | — | 79 | 83 | 76 | 82 |
| **Study 4b: Correlated Measurement Errors** | | | | |
| Measurement errors of x are correlated .50 with those of y | — | — | 48 | 45 | 48 | 47 | — | — | 3 | 2 | 10 | 2 | — | — | .39 | .38 | .42 | .40 | — | — | 66 | 66 | 63 | 65 |
| Measurement errors of x are correlated .50 with those of z | — | — | 6 | 8 | 7 | 6 | — | — | 4 | 2 | 9 | 2 | — | — | .37 | .36 | .40 | .38 | — | — | 65 | 64 | 60 | 63 |
| Measurement errors of x are intercorrelated .50 | — | — | 21 | 12 | 21 | 21 | — | — | 3 | 2 | 6 | 2 | — | — | .34 | .35 | .35 | .35 | — | — | 60 | 60 | 57 | 59 |
| **Study 4c: Misspecification of the Structural Model** | | | | |
| Correlation X with Z is 0 | — | — | 1 | 2 | 1 | 1 | — | — | 10 | 11 | 13 | 14 | — | — | .38 | .37 | .39 | .39 | — | — | 8 | 8 | 8 | 9 |
| Correlation X with Z is .20 | — | — | 8 | 9 | 8 | 9 | — | — | 10 | 10 | 14 | 13 | — | — | .40 | .39 | .42 | .41 | — | — | 18 | 21 | 17 | 22 |
| Correlation X with Z is .40 | — | — | 14 | 15 | 14 | 15 | — | — | 9 | 7 | 12 | 9 | — | — | .45 | .44 | .46 | .47 | — | — | 43 | 48 | 40 | 48 |
| Correlation X with Z is .60 | — | — | 18 | 18 | 18 | 19 | — | — | 7 | 5 | 11 | 7 | — | — | .52 | .51 | .53 | .54 | — | — | 68 | 73 | 65 | 72 |

Study 4c generates a polynomial of $X (Y = \beta_1 \times X + \beta_2 \times Z + \beta_4 \times X^2)$ and estimates the parameters with Equation 1.

Notes: Boldface indicates acceptable levels of bias: parameter bias (maximum of the moderation and main effects) ≤ 10% and standard error bias ≤ 5% (Feingold 2019; Muthén and Muthén 2002) at a sample size of 200, which is about the median in the literature review (Table 1). Reported RMSE sums the RMSE of the moderation and main effects, and reported power is the estimated power of $\beta_3$ as target moderation test. The multigroup and means methods were excluded from Studies 3 and 4a–4c on the basis of their performance in Studies 1 and 2a–c. Methods are 1: multigroup, 2: means, 3: corrected means, 4: factor scores, 5: product indicators, and 6: latent product.
smaller than .95. Its standard error bias decreases when sample size increases (e.g., bias of about 5% at a reliability of .75 and sample size of 1,500).

**Root mean square error (Panel C).** The RMSE differences are small among the unbiased methods (e.g., RMSE between .12 for the factor scores method and .14 for the product indicators method at a reliability of .85 and sample size of 200). The means method offers the best RMSE in smaller samples (≤ 500 observations). However, it is biased and should therefore not be used. The product indicators method has a high RMSE, .55 at a reliability of .75 and a sample size of 200, due to its upward standard error bias.

**Power (Panel D).** Among unbiased methods, the factor scores method has the highest power: an estimated 65% at measure reliabilities of .85 and a sample size of 200. Its power is 51% at a reliability level of .75 and 80% for reliabilities of .95. However, power differences between the corrected means method and the latent product method are only one to three percentage points across conditions. At reliabilities of .75 and a sample size of 200, the multigroup method (34% power due to discretization) and the product indicators method (37% power due to standard error bias) have lower power than the remaining methods.

**Discussion**

Study 1 raises concerns about the performance of the means, multigroup, and product indicators methods, even at reliabilities of .85, which are conventional in the literature review (Table 1) and commonly considered good (Peterson 1994). In contrast, the corrected means, factor scores, and latent product methods perform relatively well across conditions. Their parameter bias is under 10% and standard error bias is below 5% (Feingold 2019; Muthén and Muthén 2002) at a sample size of 200 (and higher). There are few differences in power and RMSE between these three methods. Main effect results offer similar conclusions (Web Appendix F).

However, the estimated power to find a small-to-medium moderation effect of .20 (about the mean in the literature review; see Table 1) at a measure reliability level of .85 (about the mean) and a sample size of 200 (about the median) is only about 65% at best. To estimate required sample sizes for 80% power based on Study 1, we follow Schoemann et al. (2014) and extract fitted probabilities from a binary probit regression of the significance of the moderation effect (1 if it is statistically significant, 0 otherwise) on an intercept, the sample size, the dummy-coded reliability, the dummy-coded method, and all interactions. The estimated required sample size is then the smallest sample for which the estimated likelihood (power) of a statistically significant moderation effect is at least 80%.

Table 6 reports the estimates. To find a moderation effect of .20 at a reliability of .85 and with 80% power, the corrected means, factor scores, and latent product methods need at least 312 observations. This requirement is almost 50% larger than the median sample size of 215 in the literature review and is only met by 33% of studies in our literature review. Thus, larger sample sizes are needed to attain sufficient power. At a high reliability of .95, slightly more than 200 observations are sufficient for 80% power. Smaller reliabilities of .75 require even larger samples (e.g., ≥450 for the latent product method). These results are in line with findings in the strategic management domain (Aguinis, Edwards, and Bradley 2017) and suggest that a substantive proportion of published moderation effects under investigation might be biased downward (because of the widespread use of the means method) and/or underpowered (because of moderation analysis in small samples).

**Study 2a: Ordered Categorical Indicators**

**Design**

Study 2a extends Study 1 by using ordered categorical indicators rather than continuous indicators. The design is 6 (method) × 7 (sample size) × 3 (number of scale points of y, x, and z: seven, five, or three). We follow Rhemtulla, Brosseau-Liard, and Savalei (2012) and use thresholds based on Z-scores that equally divide ±2.5 standard deviations from the mean to transform the continuous indicators. We focus on seven-point scales (58% of the cases in the literature review), five-point scales (13%), and three-point scales (below 1%) to explore boundary conditions. Overall, categorical indicators contain less information than continuous indicators do, but Rhemtulla, Brosseau-Liard, and Savalei (2012) find that indicators with five or more ordered categories perform similarly to continuous indicators in nonmoderation settings. Study 2a tests whether this holds for latent moderation.

**Results**

First, the bias of the multigroup and means methods increases when the number of scale points decreases. For instance, the parameter bias increases from 27% (seven-point) to 38% (three-point) for the means method (compared with 26% for continuous indicators in Study 1). Second, the factor scores method and the latent product method remain unbiased (parameter and

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2 There is a possibility that the performance of the methods differs for specific subsets of the data. For instance, methods that are more heavily parameterized (such as the product indicators method and the latent product method) might be more prone to fitting idiosyncrasies in the data (e.g., sampling error) instead of recovering the true moderation effect, which is undesirable. We conduct tenfold cross-validation (James et al. 2013, p. 181; Singh, Marinova, and Singh 2020) to examine this. We use the four foci performance criteria to compare the methods. Preferred methods should only have small differences in terms of the in-sample performance criteria with those based on tenfold cross-validation. Web Appendix F summarizes cross-validation results of Study 1 that have only small differences with the in-sample performance. This finding is encouraging and rules out overfitting. Because we find little substantive difference between the in-sample performance of the methods and the performance based on tenfold cross-validation, we do not conduct cross-validation for Studies 2a–c, 3, and 4a–c.
standard error bias below 5%) across conditions, and their RMSE and power levels are similar (e.g., RMSE of .37 for factor scores and .39 for the latent product method). However, power levels are lower than in Study 1. The latent product method has 58%, 56%, and 45% power for seven-, five-, and three-point scales at a reliability of .85 and sample size of 200, whereas it had 64% power in Study 1. Third, the corrected means and product indicators methods are biased for three-point scales (standard error bias up to 43% at a sample size of 200). However, and interestingly, the product indicators method has a standard error bias of 5% for at least five-point scales, whereas it had standard error bias of 10% at a sample size of 200 for continuous indicators (Study 1). In this simulation, categorical scales limit extreme values in the indicators, such as outliers, that are more likely to occur for continuous scales and become bigger issues as a result of indicator multiplication. In summary, although categorical indicators contain less

Figure 2. Study 1: Performance criteria for the moderation effect ($\beta_3$).
Notes: Plots visualize method parameter bias, standard error bias, root mean square error (RMSE), and power (as defined in Table 2) of the moderation effect ($\beta_3$) across sample sizes (log scale) and reliabilities of Y, X, and Z. Horizontal dashed lines indicate parameter bias, standard error bias, and RMSE of zero and power of 80%. Vertical dashed lines indicate a sample size of 200, which is about the median in the literature review (Table 1).
information than continuous ones, leading to lower power, five-point and seven-point scales perform almost equally to continuous indicators in terms of unbiasedness for the factor scores method and the latent product method. The factor scores method and the latent product method outperform the corrected means method and the product indicators method for three-point scales.

**Study 2b: Discrete Moderator**

**Design**

Study 2b extends Study 1 by focusing on a single discrete (binary) moderation indicator without measurement error (e.g., a country indicator or a manipulation dummy). This is the case for about a third of the moderation effects in the literature review. The design is 6 (method) × 7 (sample size) × 3 (reliability of Y, X, and Z: .95, .85, or .75). Here, the multigroup method can use the moderator without discretization, and we investigate how the multigroup method performs in comparison with the other methods in such a setting.

**Results**

First, the multigroup method is unbiased (bias under 2% across sample sizes) for a discrete moderator. Similarly, standard error biases are under 5% at a sample size of 200. Second, the bias of...
the means method is lower than in Study 1 but persists (about 15% at a reliability of .85) unless reliabilities are .95 (bias about 5%). Third, the parameter and standard error biases of the product indicators method are lower than in Study 1 (below 5% across reliabilities of .75 to .95 and for sample sizes of 200 and larger). These findings are due to the fact that Study 2b only has measurement error in x and y, whereas Study 1 focused on measurement error in y, x, and z. Fourth, the corrected means, factor scores, product indicators, and latent product methods are unbiased (parameter and standard error bias up to 5%). We find that RMSE (e.g., between .33 [factor scores and corrected means] and .35 [multigroup] at a reliability of .85 and sample size of 200) and power (70%–71%) are similar under the investigated conditions for the unbiased methods. In summary, the multigroup method is a well-performing alternative to the corrected means, factor scores, product indicators, and latent product methods for binary moderators without measurement error.

Study 2c: Correlation of X with Z

Design

Study 2c extends Study 1 by varying the correlation between X and Z (fixed to .20 in Study 1). Typically, X and Z are correlated in observational data, and this might impact method performance (Grewal, Cote, and Baumgartner 2004). The design is 6 (method) × 7 (sample size) × 4 (correlation of X with Z: 0, .20, .40, .60) with the correlation varying from 0 to .60, covering most of the range in the literature review (Web Appendix A).

Results

First, the bias of the moderation effect for the multigroup method decreases when the correlation between X and Z increases, but the main effects (see Web Appendix I) become more biased (up to 86% at a correlation of .60). Second, increasing the correlation between X and Z from 0 to .60 decreases the moderation bias for the means method from 27% to 21%. This finding is due to the higher reliability of product terms for correlated components (see Equation 6). Third, the corrected means, factor scores, and latent product methods are unbiased across conditions (parameter bias below 3% and standard error bias below 5%), whereas the product indicators method has standard error bias of 9%–10%. The unbiased methods have similar RMSE and power levels (e.g., RMSE between .40 [factor scores] and .42 [latent product method] at a correlation of .60 and sample size of 200). Fourth, the power of the moderation effect increases from 62% to 79% for the latent product method when the correlation between X and Z increases from 0 to .60. Thus, the increase in the power of the moderation effect due to the increased reliability of the product term trades off against the decrease in power due to multicollinearity. However, consistent with Grewal, Cote, and Baumgartner (2004), the power of the main effects decreases because of multicollinearity (from 67% to 47% for the main effects; see Web Appendix I). In summary, the corrected means, factor scores, and latent product methods are unbiased under the investigated conditions. Higher correlation between X and Z increases power to find a moderation effect but decreases power of the main effects.

Study 3: Unequal Indicator Loadings

Design

Study 3 extends Study 1 by focusing on indicators of the latent variables that differ in their loadings. Because the multigroup method and the means method are biased across conditions in Studies 1, 2a, and 2c, on which the following studies build, Studies 3 and 4a–c focus on the comparison between the remaining methods: factor scores, corrected means, product indicators, and latent product. The design is 4 (method) × 7 (sample size) with unequal indicators in all cells: \( \lambda_{x1} = 1, \lambda_{x2} = 1.5, \lambda_{x3} = .50 \) (and analogous for Z and Y). We hold indicator measurement error variances constant such that measure (composite) reliabilities are equal to those from Study 1 and to make sure that differences between equal and unequal loading conditions are not confounded with differences in measure reliability. We expect the factor scores method and the latent product method to perform best because they freely estimate loadings. The corrected means method assumes that all indicators are equally good representatives of their underlying latent factors, and Cronbach’s alpha underestimates measure reliability if this assumption is violated (McNeish and Wolf 2020). This underestimation might lead to measurement error corrections that bias the estimates upward.

Results

First, as expected, the corrected means method has biased moderation and main effect estimates, at least 20% even at large sample sizes of 1,500. Second, the factor scores and latent product
methods that freely estimate indicator loadings perform best, with parameter and standard error biases under 5% at sample sizes of 200 (and higher). Their RMSE and power are similar (e.g., 69% power of the factor scores method and 68% power of the latent product method at a sample size of 200). Third, the product indicators method has a low parameter bias, as do the factor scores method and the latent product method, but it has a higher standard error bias (e.g., 12% at a sample size of 200). In summary, the factor scores method and the latent product method perform best for unequal indicator loadings.

**Study 4a: Nonnormally Distributed Indicators**

**Design**

Studies 4a–c investigate situations where model assumptions of all focal methods are violated, unlike Study 3, where assumptions are violated for the corrected means method only. Study 4a focuses on nonnormality distributed x and z, which is common when measuring constructs as customer satisfaction (Peterson and Wilson 1992). The design is 4 (method) × 7 (sample size) × 2 (skewness and excess kurtosis for X and Z: 1 and 2, or 3 and 10). Skewness and excess kurtosis are conventional metrics of nonnormality. Both are zero for normally distributed variables (Oliveira, Oliveira, and Seijas-Macias 2016). Because we could not determine skewness and excess kurtosis in our literature review, we use about the 75th and 95th percentiles from a recent existing review in psychology (Cain, Zhang, and Yuan 2017, p. 1720). The procedure described by Vale and Maurelli (1983) generates nonnormal latent variables X and Z that are reflected in nonnormal indicators. Previous research concluded that nonzero skewness and excess kurtosis in variables lead to overestimated zero-order correlations (Bishara and Hittner 2015) but underestimated standard errors (Finch, West, and MacKinnon 1997). Yet, there might be differences between methods. Biased reliability estimates due to nonnormality can bias the corrected means method (Sheng and Sheng 2012). The product indicators method was found to be robust for different latent variable distributions (Marsh, Wen, and Hau 2004) although taking multiple indicator products might also exacerbate bias due to nonnormality. The latent product method does not use (algebraic) multiplications of indicators, so it might perform better, but severe nonnormality can still hamper the ability of the mixture distribution to approximate the indicator distribution (Klein and Moosbrugger 2000).

**Results**

First, all methods are biased (up to 19% for the corrected means method) in presence of severe nonnormality in x and z (i.e., skewness of X and Z is 3 and excess kurtosis is 10). One exception is the product indicators method with 5% parameter bias at a sample size of 200. Second, standard errors of all methods are also biased, including those of the product indicators method (standard error bias of 32%). Third, for moderately nonnormally distributed indicators (i.e., skewness of X and Z is 1 and excess kurtosis is 2), the factor scores method and the latent product method have biases under 5%. Their RMSE and power levels are similar (e.g., RMSE .35–.38 at a sample size of 200). In summary, the expectation of severe nonnormally distributed indicators with skewness and excess kurtosis might call for the product indicators method even though its statistical conclusion validity might be questionable because of biased standard errors.

**Study 4b: Correlated Measurement Errors**

**Design**

Study 4b focuses on another type of misspecification: correlated measurement errors. The design is 4 (method) × 7 (sample size) × 3 (measurement error correlation: x with y, x with z, or x with x). Correlated measurement errors can occur because of omitted variables in the measurement model, such as method factors or response tendencies (Baumgartner and Weijters 2017). We focus on three types of measurement error correlations. First, we generate error correlations between indicators of x and y (denoted as “x with y”). Evans (1985) and Siemsen, Roth, and Oliveira (2010) showed in the context of the means method that measurement error correlations between x and y do not bias moderation effects upward but can bias them downward depending on the magnitude of measurement error correlation. However, it is unclear whether these results hold for the main effects, for the other methods, and for other measurement error correlations. Hereinafter, the design also includes measurement error correlation between moderation indicators x and z and for indicators of X with other indicators of X (denoted as “x with x”). For brevity, we do not focus on measurement error correlations of z with y (analogous to x with y) and z with z (analogous to x with x). The measurement error correlation in all cells is .50. To generate the correlated measurement errors for x with y (analogous for x with z), we correlate indicator x1 with y1, x2 with y2, and x3 with y3. Measurement error correlations of x with x intercorrelate all three indicators of X.

**Results**

First, measurement error correlations of .50 between x and y bias the main effect estimate of X up to about 50% for all methods (see Web Appendix L for details) even though the moderation effect is unbiased (under 5%). This result extends what was previously found for the means method (Evans 1985; Siemsen, Roth, and Oliveira 2010). Second, under the investigated conditions, measurement error correlations of x with z yield parameter biases under 10% for the moderation and main effects across methods, much less than for measurement error correlations between x and y. However, the standard error bias of the product indicators method is 9%, whereas the standard error bias of the corrected means, factor scores, and latent product methods is 2%–4%. Third, measurement error correlations of x with x also severely bias the moderation and main effects of the corrected means method, product indicators method, and latent product method for about 21%. However, the bias is 12%
(about 9% less) for the factor scores method. One reason for this result might be that the two-step estimation of the factor scores method, compared with one-step or simultaneous estimation of the latent product method, is more robust to misspecification in the measurement model (Devlieger and Rosseel 2017; Rosseel 2020; Smid and Rosseel 2020). Thus, under the investigated conditions, correlated measurement error biases all methods. The bias is most severe for measurement error correlations of predictors with outcomes (e.g., \(x\) with \(y\)).

**Study 4c: Misspecification of the Structural Model**

**Design**

Study 4c focuses on misspecification of the structural model. The design is 4 (method) \(\times 7\) (sample size) \(\times 4\) (correlation of \(X\) with \(Z\)) = 0, .20, .40, .60). It generates a true U-shaped effect of \(X\) on \(Y\) (i.e., \(Y = \beta_1 \times X + \beta_2 \times Z + \beta_3 \times X^2\)) and uses the structural model in Equation 1 for estimation. Because moderation product terms and squared terms are generally correlated because of their common lower-order components if they are not manipulated (Ganzach 1997), the design varies the correlation between \(X\) and \(Z\). Although we expect little differences between the methods, it is difficult to quantify bias and resulting Type I error analytically.

**Results**

First, when \(X\) and \(Z\) are uncorrelated, we find that the methods yield unbiased (\(\leq 2\%\)) moderation effects. Bias for all methods is just under 10% when \(X\) and \(Z\) are correlated at .20. Second, when the correlation between \(X\) and \(Z\) increases, the bias due to misspecification increases, for instance to 19% for the latent product method and at a correlation of \(X\) with \(Z\) of .60 and a sample size of 200. Third, standard errors of all methods are biased between 5% and 15% across conditions, even at large sample sizes of 1,500. Fourth, all methods have Type I error \(\geq 5\%\) across conditions, about 20% at a correlation of .20 and a sample size of 200, which further increases if the correlation between \(X\) and \(Z\) or the sample size increases.

**General Discussion**

We compared six methods for latent moderation analysis and provide several recommendations. First, the choice between five of the six methods is at the researcher’s discretion when reliabilities of the moderating variables approach 1. Although the multigroup method is biased by over 20% when the indicators of the moderator are continuous, the parameter bias of the corrected means, factor scores, product indicators, and latent product methods across sample sizes is under 2%, and the standard error bias is under 5% when the reliability of \(Y\), \(X\), and \(Z\) was high at .95 (Study 1). The parameter bias of the means method is then 8% (and standard error bias is 3%), which might be acceptable (Table 2). Furthermore, RMSE and power differences between methods were small. The closer the reliabilities of the moderating variables are to 1, the more similar the performance of five of the six methods becomes.

Yet, reliabilities of moderation variables approaching 1 are rare in practice: the mean reliability in the literature review was .87 (Table 1), and only 13% of moderation tests had reliabilities of the moderation variables \(\geq .95\). Thus, our findings and recommendation are in contrast with the use of the means method in 95% of the literature review (Table 1). It is well known that ignoring measurement error can bias parameter estimates (Grewal, Cote, and Baumgartner 2004; Spearman 1904; Wooldridge 2015). We show the bias of the means method once more, and our Monte Carlo studies quantify it in the latent moderation context: the moderation effect bias of the means method is about 40% and 25% respectively at reliabilities of .75 and .85.

Second, the factor scores method and the latent product method are recommended across most investigated conditions (Table 5). When indicators are continuous (Study 1) or seven-, five-, or three-point ordered categorical (Study 2a), or when the moderator is binary (Study 2b), and across reliabilities between .75 and .95 (Studies 1 and 2b), the factor scores method and the latent product method have parameter and standard error biases \(\leq 5\%\). The bias remains small when the correlation between moderation variables increases from 0 to .60 (Study 2c) and for unequal indicator loadings (Study 3). We conclude from the small RMSE and power differences across the conditions of our studies that the choice between the factor scores method and the latent product method is mostly at the researcher’s discretion. Method accessibility can then be relevant. Factor scores are available in most general statistical software packages. The latent product method is to our knowledge currently only available in Mplus (Muthén and Muthén 2019) and in an R package (Umbach et al. 2017). A follow-up study in Web Appendix N compares both latent product implementations and recommends Mplus in terms of performance, computation time, and the range of possible models that can be estimated. One key researcher decision in the use of the latent product method is the number of mixture components. A follow-up study in Web Appendix O shows that the default in Mplus is adequate to estimate a single moderation effect (Klein and Moosbrugger 2000).

When the factor scores method is used, decisions need to be made about the type of measurement model and the factor scores estimation method. We draw from Skrondal and Laake (2001) and Devlieger, Mayer, and Rosseel (2016) and our own analyses to recommend the following two-step factor scores (TSFS) method for latent moderation analysis.

**Step 1** is to conduct a confirmatory factor analysis with the outcome (\(Y\)) as a single factor (1-CFA) and extract Bartlett factor scores. The 1-CFA uses optimal indicator weighting, and Bartlett factor scores account for measurement error in the outcome. Then, conduct a separate confirmatory factor analysis for the predictors (\(X\) and \(Z\)) simultaneously with two factors that are allowed to correlate (2-CFA, no cross-loadings), and extract regression factor scores to assure optimal indicator weighting and account for measurement error in predictors.
Step 2 is to compute the product term from the factor scores of the predictors (multiply) and estimate moderation and main effects with the target regression or path model. We demonstrate that this TSFS method for latent moderation analysis performs well across the examined range of conditions, and about as well as the latent product method that estimates the measurement and structural models simultaneously. Web Appendix P contains a follow-up simulation study that empirically examines the harm of using different factor scores methods than those recommended here.

Third, the multigroup, corrected means, and product indicators methods are best reserved for specific settings. The multigroup method can be used for discrete moderators (bias less than 5%), although the corrected means, factor scores, product indicators, and latent product methods perform similarly. Other researchers have found the corrected means method to perform well and similarly to the latent product method for single indicators (Hsiao, Kwok, and Lai 2021). In that case, the factor scores method cannot be used. Still, if a single indicator that contains measurement error is available, it becomes more difficult to estimate unreliability and hence to account for it, compared with multi-indicator measures, for which reliability estimators are readily available (Kamakura 2015). We refer to Pieters (2017, pp. 699–700) and the references therein for guidance. We identify one setting in which the product indicators method outperforms the factor scores method and the latent product method. The product indicators method had an estimated parameter bias of about 5% (parameter bias of 14% for the factor scores method and the latent product method) when the moderation variables had a skewness of 3 and excess kurtosis of 10 and at a sample size of 200 (Study 4a). Yet, standard errors of the product indicators method, as well as those of the other methods, remain biased (e.g., 32% standard error bias for the latent product method at a sample size of 200), which can harm statistical conclusion validity. Overall, these recommendations should provide actionable guidelines for method use. Web Appendix B provides an overview of sample code for method implementation in SPSS, Stata, R, and Mplus, made available on OSF (https://osf.io/py7jx/).

In some situations the corrected means, factor scores, product indicators, and latent product methods all perform poorly. First, although we showed that correlations between (latent variables) X and Z up to .60 have a negligible effect on the bias of these methods (Study 2c), they can be biased when measurement errors of individual indicators (e.g., x and z) are correlated, independent of the correlation between X and Z. This may occur, for instance, when indicators of X and/or Z are regular and reversed items. Then, the measurement model needs to be adapted (e.g., Baumgartner and Weijters 2017; Weijters, Baumgartner, and Schillewaert 2013), such as by introducing a method factor or having specific errors correlate, before applying the methods that we have compared here. Second, if the true effect of X on Y is U-shaped (polynomial) but not specified (Hutchinson, Kamakura, and Lynch 2000), not only the means method (Ganzach 1997) but all methods perform poorly (Study 4c). That is, if the data-generating process is a U-shaped effect of X on Y, using a specification of the moderation without the polynomial X^2 leads to Type I errors for all methods. Then, a nonzero moderation effect between X and Z might be observed even though none exists in the data, a situation that can avoided by examining moderation and curvilinear effects simultaneously (Ganzach 1997).

Among our findings, the small differences in performance between the TSFS method and the latent product method across the focal conditions are noteworthy. One might have expected that joint estimation of the measurement and structural models by the latent product method would empirically perform better than the TSFS method. Recent research in the nonmoderation context has drawn attention to the role of two-step versus conventional simultaneous estimation of latent variable models (Devlieger and Rosseel 2017; Rosseel 2020; Smid and Rosseel 2020). Conceptually, the TSFS method matches the estimation of the measurement and structural models as a combination of two separate models (Anderson and Gerbing 1988). Empirically, one advantage of two-step estimation is that measurement model misspecification might lead to less structural model bias, or vice versa (Devlieger and Rosseel 2017). Our Study 2c showed this to be the case in the context of within-construct correlated measurement errors, although the reduction in bias of the factor scores method compared with the latent product method was a modest 9% under the investigated conditions. Moreover, two-step methods might have fewer convergence issues or ineligible solutions than simultaneous estimation methods do (Smid and Rosseel 2020). Follow-up analyses of our Study 1 show that although nonconvergence was rare, all replications converged for the TSFS method and the corrected means method (both two-step methods). In contrast, 2.4% of replications for the product indicators method and less than 1% of replications for the latent product method (both one-step methods) did not converge. Among nonconverging replications, small sample sizes of 100 or 150 (about 84%) and low reliabilities of .75 (about 86% of nonconverging replications) were most common, which might support the use of two-step estimation to avoid convergence issues of simultaneous estimation in such settings (Rosseel 2020). In summary, the TSFS method for latent moderation analysis is accessible, and its estimates have a small bias with low variance across a large range of conditions.

With this foundation, our study opens several avenues for further research. First, one might investigate the performance of Bayesian estimation, which might do well in small samples and facilitates the incorporation of prior information, potentially resulting in more precise estimates and moderation tests with higher power. Asparouhov and Muthén (2021) conduct simulation studies for the latent product method. Second, although this research studied both random and correlated indicator

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3 A Web of Science citation analysis of Anderson and Gerbing (1988), an early study advocating for two-step estimation of measurement and structural models, showed that the total number of citations per year increased from 149 in 2000 to 244 in 2005, 614 in 2010, 1,140 in 2015, and 2,143 in 2020. This increase might signal a more general trend of using two-step estimation for latent variable models.
measurement error, it does not focus on methods to account for correlated measurement errors. If correlated measurement errors are expected, the latent product method might be preferred over the factor scores method because it uses separate factor analyses for predictors and outcomes in which error correlations between predictors and outcomes cannot be accounted for. The latent product method estimates the measurement and structural models simultaneously. Further research might investigate this. We refer to Baumgartner and Weijters (2017) for an overview of models to account for correlated measurement errors in a nonmoderation setting. Third, although the Monte Carlo simulations study a variety of conditions, including settings that violate assumptions, the simulations can be extended further. For instance, the question remains how the methods perform for multilevel or multitime data and fixed or random effects models. Similarly, method performance in the case of (latent) instrumental variables can be assessed.

In summary, it is hard to justify the continued use of the means method for latent moderation analysis unless measurement reliabilities approach 1. Researchers are well advised to apply other methods for latent moderation analysis, such as the TSFS method and the latent product method. We hope that our recommendations improve moderation theory testing and help marketing researchers planning latent moderation studies.

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References
Aguinis, Herman, Jeffrey R. Edwards, and Kyle J. Bradley (2017). “Improving Our Understanding of Moderation and Mediation in Strategic Management Research,” Organizational Research Methods, 20 (4), 665–85.

Anderson, James C. and David W. Gerbing (1988), “Structural Equation Modeling in Practice: A Review and Recommended Two-Step Approach,” Psychological Bulletin, 103 (3), 411–23.

Asparouhov, Tihomir and Bengt O. Muthén (2021), “Bayesian Estimation of Single and Multilevel Models with Latent Variable Interactions,” Structural Equation Modeling: A Multidisciplinary Journal, 28 (2), 314–28.

Atasoy, Ozgun and Carey K. Morewedge (2017), “Digital Goods Are Valued Less Than Physical Goods,” Journal of Consumer Research, 44 (6), 1343–57.

Auh, Seigyoung, Bulent Menguc, Constantine S. Katsikeas, and Yeon Sung Jung (2019), “When Does Customer Participation Matter? An Empirical Investigation of the Role of Customer Empowerment in the Customer Participation–Performance Link,” Journal of Marketing Research, 56 (6), 1012–33.

Baumgartner, Hans and Bert Weijters (2017), “Measurement Models for Marketing Constructs,” in Handbook of Marketing Decision Models, Berend Wierenga and Ralf van der Lans, eds. Cham, Switzerland: Springer, 259–95.

Bishara, Anthony J. and James B. Hittner (2015), “Reducing Bias and Error in the Correlation Coefficient Due to Nonnormality,” Educational and Psychological Measurement, 75 (5), 785–804.

Bollen, Kenneth A. (1989), Structural Equations with Latent Variables. New York: John Wiley & Sons.

Busemeyer, Jerome R. and Lawrence E. Jones (1983), “Analysis of Multiplicative Combination Rules When the Causal Variables Are Measured with Error,” Psychological Bulletin, 93 (3), 549–62.

Cain, Meghan K., Zhiyong Zhang, and Ke-Hai Yuan (2017), “Univariate and Multivariate Skewness and Kurtosis for Measuring Nonnormality: Prevalence, Influence and Estimation,” Behavior Research Methods, 49 (5), 1716–35.

Charles, Eric P. (2005), “The Correction for Attenuation Due to Measurement Error: Clarifying Concepts and Creating Confidence Sets,” Psychological Methods, 10 (2), 206–26.

Cohen, Jacob (1988), Statistical Power Analysis for the Behavioral Sciences, 2nd ed. Hillsdale, NJ: Lawrence Erlbaum Associates.

Cohen, Jacob, Patricia Cohen, Stephen G. West, and Leona S. Aiken (2003), Applied Multiple Regression/Correlation Analysis for the Behavioral Sciences, 3rd ed. Mahwah, NJ: Lawrence Erlbaum Associates.

Dempster, A.P., N.M. Laird, and D.B. Rubin (1977), “Maximum Likelihood from Incomplete Data via the EM Algorithm,” Journal of the Royal Statistical Society: Series B (Methodological), 39 (1), 1–38.

Devlieger, Ines, Axel Mayer, and Yves Rosseel (2016), “Hypothesis Testing Using Factor Score Regression: A Comparison of Four Methods,” Educational and Psychological Measurement, 76 (5), 741–70.

Devlieger, Ines and Yves Rosseel (2017), “Factor Score Path Analysis,” Methodology, 13 (1), 31–38.

Dimitrak, Polina, Karin Schermelleh-Engel, Augustin Kelava, and Helfried Moosbrugger (2007), “Challenges in Nonlinear Structural Equation Modeling,” Methodology: European Journal of Research Methods for the Behavioral and Social Sciences, 3 (3), 100–114.

Eisend, Martin (2015), “Have We Progressed Marketing Knowledge? A Meta-Meta-Analysis of Effect Sizes in Marketing Research,” Journal of Marketing, 79 (3), 23–40.
Evans, Martin G. (1985), “A Monte Carlo Study of the Effects of Correlated Method Variance in Moderated Multiple Regression Analysis,” Organizational Behavior and Human Decision Processes, 36 (3), 305–23.

Feingold, Alan (2019), “Time-Varying Effect Sizes for Quadratic Growth Models in Multilevel and Latent Growth Modeling,” Structural Equation Modeling: A Multidisciplinary Journal, 26 (3), 418–29.

Finch, John F., Stephen G. West, and David P. MacKinnon (1997), “Effects of Sample Size and Nonnormality on the Estimation of Mediated Effects in Latent Variable Models,” Structural Equation Modeling: A Multidisciplinary Journal, 4 (2), 87–107.

Foldnes, Njål and Knut Arne Hagtvet (2014), “The Choice of Product Indicators in Latent Variable Interaction Models: Post Hoc Analyses,” Psychological Methods, 19 (3), 444–57.

Ganzach, Yoav (1997), “Misleading Interaction and Curvilinear Terms,” Psychological Methods, 2 (3), 235–47.

Germann, Frank, Peter Ebbes, and Rajdeep Grewal (2015), “The Chief Marketing Officer Matters!” Journal of Marketing, 79 (3), 1–22.

Grewal, Rajdeep, Joseph A. Cote, and Hans Baumgartner (2004), “Multicollinearity and Measurement Error in Structural Equation Models: Implications for Theory Testing,” Marketing Science, 23 (4), 519–29.

Hallquist, Michael N. and Joshua F. Wiley (2018), “MplusAutomation: An R Package for Facilitating Large-Scale Latent Variable Analyses in Mplus,” Structural Equation Modeling: A Multidisciplinary Journal, 25 (4), 621–38.

Hsiao, Yu-Yu, Oi-Man Kwok, and Mark H.C. Lai (2021), “Modeling Measurement Errors of the Exogenous Composites from Congeneric Measures in Interaction Models,” Structural Equation Modeling: A Multidisciplinary Journal, 28 (2), 250–60.

Hunter, John E. and Frank L. Schmidt (2004), Methods of Meta-Analysis: Correcting Error and Bias in Research Findings, 2nd ed. Thousand Oaks, CA: SAGE Publications.

Hutchinson, J. Wesley, Wagner A. Kamakura, and John G. Lynch (2000), “Unobserved Heterogeneity as an Alternative Explanation for ‘Reversal’ Effects in Behavioral Research,” Journal of Consumer Research, 27 (3), 324–44.

Irwin, Julie R. and Gary H. McClelland (2001), “Misleading Heuristics and Moderated Multiple Regression Models,” Journal of Marketing Research, 38 (1), 100–109.

Irwin, Julie R. and Gary H. McClelland (2003), “Negative Consequences of Dichotomizing Continuous Predictor Variables,” Journal of Marketing Research, 40 (3), 366–71.

James, Gareth, Daniela Witten, Trevor Hastie, and Robert Tibshirani (2013), An Introduction to Statistical Learning, 1st ed. New York: Springer.

Jöreskog, Karl G. and Fan Yang (1996), “Nonlinear Structural Equation Models: The Kenny-Judd Model with Interaction Effects,” in Advanced Structural Equation Modeling: Issues and Techniques, G. Marcoulides and R. Schumacker, eds. New York: Psychology Press, 57–88.

Kamakura, Wagner A. (2015), “Measure Twice and Cut Once: The Carpenter’s Rule Still Applies,” Marketing Letters, 26 (3), 237–43.

Kenny, David A. and Charles M. Judd (1984), “Estimating the Nonlinear and Interactive Effects of Latent Variables,” Psychological Bulletin, 96 (1), 201–10.

Klein, Andreas and Helfried Moosbrugger (2000), “Maximum Likelihood Estimation of Latent Interaction Effects with the LMS Method,” Psychometrika, 65 (4), 457–74.

Lastovicka, John L. and Kanchana Thomodaran (1991), “Common Factor Score Estimates in Multiple Regression Problems,” Journal of Marketing Research, 28 (1), 105–12.

Lin, Guan-Chyun, Zhonglin Wen, Herbert W. Marsh, and Huey-Shyan Lin (2010), “Structural Equation Models of Latent Interactions: Clarification of Orthogonalizing and Double-Mean-Centering Strategies,” Structural Equation Modeling: A Multidisciplinary Journal, 17 (3), 374–91.

Marsh, Herbert W., Zhonglin Wen, and Kit-Tai Hau (2004), “Structural Equation Models of Latent Interactions: Evaluation of Alternative Estimation Strategies and Indicator Construction,” Psychological Methods, 9 (3), 275–300.

McClelland, Gary H., Julie R. Irwin, David Disatnik, and Liron Sivan (2017), “Multicollinearity Is a Red Herring in the Search for Moderator Variables: A Guide to Interpreting Moderated Multiple Regression Models and a Critique of Iacobucci, Schneider, Popovich, and Bakamitos (2016),” Behavior Research Methods, 49 (1), 394–402.

McNeish, Daniel and Melissa Gordon Wolf (2020), “Thinking Twice About Sum Scores,” Behavior Research Methods, 52 (6), 2287–305.

Moosbrugger, Helfried, Karin Schermelleh-Engel, and Andreas Klein (1997), “Methodological Problems of Estimating Latent Interaction Effects,” Methods of Psychological Research Online, 2 (2), 95–111.

Muthén, Linda K. and Bengt O. Muthén (2002), “How to Use a Monte Carlo Study to Decide on Sample Size and Determine Power,” Structural Equation Modeling: A Multidisciplinary Journal, 9 (4), 599–620.

Muthén, Linda K. and Bengt O. Muthén (2019), Mplus User’s Guide, 8th ed. Los Angeles: Muthén & Muthén.

Oliveira, Amílcar, Teresa Oliveira, and Antonio Seijas-Macías (2016), “Evaluation of Kurtosis into the Product of Two Normally Distributed Variables,” AIP Conference Proceedings, 1738 (1), 1–4.

Peterson, Robert A. (1994), “A Meta-Analysis of Cronbach’s Coefficient Alpha,” Journal of Consumer Research, 21 (2), 381–91.

Peterson, Robert A. and William R. Wilson (1992), “Measuring Customer Satisfaction: Fact and Artifact,” Journal of the Academy of Marketing Science, 20 (1), 61–71.

Pieters, Rik (2017), “Meaningful Mediation Analysis: Plausible Causal Inference and Informative Communication,” Journal of Consumer Research, 44 (3), 692–716.

R Core Team (2020), R: A Language and Environment for Statistical Computing. Vienna, Austria: R Foundation for Statistical Computing.

Rhemtulla, Mijke, Patricia É Brousseau-Liard, and Victoria Savalei (2012), “When Can Categorical Variables Be Treated as Continuous? A Comparison of Robust Continuous and Categorical SEM Estimation Methods Under Suboptimal Conditions,” Psychological Methods, 17 (3), 354–73.

Rosseel, Yves (2012), “lavaan: An R Package for Structural Equation Modeling,” Journal of Statistical Software, 48 (2), 1–36.

Rosseel, Yves (2020), “Small Sample Solutions for Structural Equation Modeling,” in Small Sample Solutions: A Guide for
Schoemann, Alexander M., Patrick Miller, Sunthud Pornprasertmanit, and Wei Wu (2014), “Using Monte Carlo Simulations to Determine Power and Sample Size for Planned Missing Designs,” International Journal of Behavioral Development, 38 (5), 471–79.

Sheng, Yanyan and Zhaohui Sheng (2012), “Is Coefficient Alpha Robust to Non-normal Data?” Frontiers in Psychology, 3 (34), 1–13.

Siemsen, Enno, Aleda Roth, and Pedro Oliveira (2010), “Common Method Bias in Regression Models with Linear, Quadratic, and Interaction Effects,” Organizational Research Methods, 13 (3), 456–76.

Singh, Sunil K., Detelina Marinova, and Jagdip Singh (2020), “Business-to-Business E-Negotiations and Influence Tactics,” Journal of Marketing, 84 (2), 47–68.

Skrondal, Anders (2000), “Design and Analysis of Monte Carlo Experiments: Attacking the Conventional Wisdom,” Multivariate Behavioral Research, 35 (2), 137–67.

Skrondal, Anders and Petter Laake (2001), “Regression Among Factor Scores,” Psychometrika, 66 (4), 563–75.

Smid, Sanne C. and Yves Rosseel (2020), “SEM with Small Samples: Two-Step Modeling and Factor Score Regression Versus Bayesian Estimation with Informative Priors,” in Small Sample Solutions: A Guide for Applied Researchers and Practitioners, Rens Van de Schoot and Milica Miočević, eds. New York: Routledge, 239–54.

Spearman, C. (1904), “The Proof and Measurement of Association Between Two Things,” The American Journal of Psychology, 15 (1), 72–101.

Umbach, Nora, Katharina Naumann, Holger Brandt, and Augustin Kelava (2017), “Fitting Nonlinear Structural Equation Models in R with Package nsem,” Journal of Statistical Software, 77 (1), 1–20.

Vale, C. David and Vincent A. Maurelli (1983), “Simulating Multivariate Nonnormal Distributions,” Psychometrika, 48 (3), 465–71.

Van Smeden, Maarten, Timothy L. Lash, and Rolf H.H. Groenwold (2019), “Reflection on Modern Methods: Five Myths About Measurement Error in Epidemiological Research,” International Journal of Epidemiology, 49 (1), 338–47.

Weijters, Bert, Hans Baumgartner, and Niels Schillewaert (2013), “Reversed Item Bias: An Integrative Model,” Psychological Methods, 18 (3), 320–34.

Wooldridge, Jeffrey M. (2015), Introductory Econometrics: A Modern Approach, 6th ed. Boston: Cengage Learning.

Yuan, Ke-Hai, Ying Cheng, and Wei Zhang (2010), “Determinants of Standard Errors of MLEs in Confirmatory Factor Analysis,” Psychometrika, 75 (4), 633–48.