Association between the Functional Polymorphism of Vascular Endothelial Growth Factor Gene and Breast Cancer: A Meta-Analysis

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Abstract

The vascular endothelial growth factor (VEGF) gene single-nucleotide polymorphism involved in the regulation of the protein levels has been implicated in breast cancer. However, the published studies have produced contentious and controversial results. Herein, we performed a meta-analysis (from January to October 2013) to further evaluate the association between +936 C/T polymorphism and the risk of breast cancer. By searching the EMBASE, PubMed, and Web of Science databases, we identified a total of 12 case-control studies with 8,979 cancer patients and 9,180 healthy controls. The strength of the association was assessed using Odds Ratios (ORs) with 95% Confidence Intervals (CI). We found no evidence indicating that the allelic model or the genotype models of +936 C/T polymorphism were associated with the risk of breast cancer in total population (OR CC vs. TT = 1.01, 95% CI = 0.96-1.06, P = 1.00; OR CC + CT vs. TT = 1.00, 95% CI = 0.96-1.05, P = 1.00; OR CC vs. CT + TT = 1.02, 95% CI = 0.98-1.07, P = 0.94; OR allele C vs. allele T = 1.01, 95% CI = 0.98-1.04, P = 0.99; OR CT vs. TT = 1.01, 95% CI = 0.93-1.09, P = 1.00). Such lack of association with breast cancer was also observed in subgroup analyses according to ethnicity as well as in the analysis by source of controls. In conclusion, this meta-analysis suggests that the functionally important +936 C/T polymorphism may not be associated with breast cancer risk. Larger well-designed studies with gene-to-gene and gene-to-environment interactions are clearly required to validate the results further.

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Keywords

● Vascular endothelial growth factor A
● Breast neoplasms
● Polymorphism
● Risk

Introduction

Angiogenesis is critical for the growth and metastasis of invasive tumors and constitutes an important component in the suppression of cancer formation.1 The process of transporting excess nutrients, producing some risk factors, and forming tumor blood vessels and a route for tumor cell egress induces tumor aggression, growth, and dissemination.2,3 Vascular endothelial growth factor (VEGF) acts as an angiogenic inducer that is an endothelial cell-specific mitogen, and as a mediator of vascular permeability, playing a central role in the regulation of this process. There has been much speculation that inhibition of VEGF activity may prevent tumor initiation, progression,
VEGF polymorphism and the risk of breast cancer in EMBASE (http://www.embase.com/), PubMed (http://www.ncbi.nlm.nih.gov/pubmed), and Web of Science (http://newisiknowledge.com) without any language restriction, using a combination of the following search terms: “vascular endothelial growth factor” or “VEGF”, “+936 C/T” or “rs3025039”, “genotypes” or “variant” or “variants” or “polyorphism” or “polymorphisms”, and “breast cancer” (the last search updated in June 2013). We also hand-searched the reference lists cited in genetic association studies, narrative reviews and meta-analyses to further identify the relevant publications that may have missed in the electronic search. We additionally carried out hand searches of the journals known to publish SNP-breast cancer articles, including International Journal of Cancer, Medical Oncology, Breast Cancer Research, Cancer Letters, Clinical Cancer Research, Cancer Epidemiology, Biomarker & Prevention, Breast Cancer Research Treatment, et al. All titles, abstracts, and full-texts were carefully reviewed to check if there were usable data reported in the original articles. When the required data were absent, corresponding authors were contacted via e-mail. However, no response was received by this process.

Inclusion and Exclusion Criteria
The following inclusion criteria were defined for the eligible studies in this meta-analysis: (a) the study must be published as a full-text paper with a case-control or cohort design; (b) the authors must have investigated the association between +936 C/T polymorphism and the risk of breast cancer; (c) the study must provide sufficient genetic data that could help to calculate an odds ratio (OR) and its 95% confidence intervals (CI); (d) the publication must be released online before the literature search was completed. We excluded the studies if; only breast cancer patients were investigated, information on genotype count was not detailed, or no responses were received from major authors. We also did not consider editorials, case reports, systematic reviews, and comment letters. In case of studies containing overlapped subjects, we selected the largest study with accessible genotype data.

Data Extraction
In order to maximize the reliability of data analyzed in this work, the data extraction was done by two independent reviewers (J Li and Yy Ding) based on a consensus reached on all items. The following characteristics were collected from each eligible article: first author’s surname, journal and year of publication, country where
the study was carried out, ethnicity or racial
descent (categorized as Caucasian or Asian),
source of controls (categorized as population-
or hospital-based study), baseline characteristics
of the breast cancer patients and the disease-
free individuals (e.g. mean age, smoking status,
family history, body mass index, menopausal
status, tumor size, regional lymph node
metastasis, distant metastasis, histologic grade
wherever available), assays utilized to genotype
the polymorphism of interest, total cases and
controls, genotype frequencies between cases
and controls, matching status, adjusted factors
and study design (categorized as retrospective
or prospective). Accuracy of the extracted
information was examined by crosscheck.
Discrepancies, if any, were resolved through
discussion with a senior reviewer (Yj Ju).

Quality Assessment
The quality of included studies was assessed
independently by the same two reviewers (J L.
Li and Yy Ding) according to the P values of
Hardy-Weinberg Equilibrium (HWE) in control
groups.32 We defined the studies whose genotype
distribution in controls was consistent with HWE
(P>0.10) as high-quality studies, and those
inconsistent with HWE as (P≤0.10) low-quality
studies. We then conducted subgroup analyses
according to genotype (high quality, low quality)
to confirm whether deviation from HWE had an
influence on the association being investigated.

Statistical Methods
For the controls of each study, HWE was
examined by the chi-square goodness-of-fit test
with P<0.05 representing a significant deviation.
The OR and its 95% CI were calculated to assess
the strength of the association between +936 C/T
polymorphism and breast cancer risk. The pooled
ORs were performed for allelic model: allele C
vs. allele T, and genotype models: CC vs. TT
(homozygote), CC+CT vs. TT (dominant), CC vs.
CT+TT (recessive), and CT vs. TT (heterozygote).
The Z test was applied to determine the statistical
significance of summary ORs, and P<0.05 was
deemed statistically significant.

Heterogeneity assumption was evaluated by
a chi-square based Q test. The significant level
was defined at P<0.10. In addition, the I^2 index
indicating the proportion of inter-study variance
attributable to heterogeneity rather than chance
was used to measure the degree of inconsistency
across studies. I^2 statistics range from 0 and
100%, with higher proportions representing
larger heterogeneity: 0-25% low heterogeneity,
25%-50% moderate heterogeneity, 50%-75%
large heterogeneity, 75%-100% extremely large
heterogeneity. The fixed effect model derived
from the Mantel-Haenszel method33 was used
to pool the ORs when P>0.10 or I^2 statistics less
than 50%; otherwise, the random effect model
described by Darlington34 was more appropriate.
In addition, subgroup analyses according to
ethnicity and source of controls were performed
to detect the potential heterogeneity further.

The influence of each individual study on the
summary ORs was evaluated by excluding the
single studies in turn from the pooled estimate.
Publication bias was examined visually by funnel-
inverted plots in which the standard error of log
(OR) of each study was plotted against its log
(OR). Egger’s test was utilized to check the
asymmetry of funnel plots on the natural logarithm
scale of the ORs.35 An asymmetric plot and the P
values of Egger’s linear regression test smaller
than 0.10 suggested obvious publication bias.

All statistical data were analyzed with Stata
software (Version 12.0, Stata Corporation,
College Station, Texas, USA). All tests were two-
sided with a statistically significant level at P<0.1.

Results
Selection and Characteristics of Eligible Studies
As depicted in figure 1, the computer-based
search of EMBASE, PubMed, and Web of
Science databases identified a total of 12 eligible
studies15-26 for this meta-analysis, involving
8,979 breast cancer cases and 9,180 cancer-
free controls. Although we primarily obtained
697 records, the initial removal of obviously
irrelevant records and duplicates resulted in 27
studies for further evaluation. We subsequently
deleted 15 studies due to various reasons,
including research on other SNPs at the same
locus, lack of data for calculation of ORs, case-
case study, and comment letters.36-50 The main
characteristics of the included studies are
summarized in table 1. In terms of ethnicity, ten
studies were based on Caucasian subjects and
two on Asian subjects. The study by Jin et al.
contained three sub-study groups from Poland,
Germany, and Sweden respectively, and the
three populations were taken as independent
studies, which were merged into Caucasians in
overall and subgroup meta-analysis. The quality
of most included studies (83.3%) was high.
Genotyping methods employed to determine
VEGF SNP +936 C/T varied widely across
studies, with the PCR-restriction fragment
length polymorphism (PCR-RFLP) being most
commonly used, followed by TaqMan and
direct sequencing. The genotype distribution in
controls of all studies was consistent with HWE
except for the studies by Lin et al. and Luo et al.
Meta-Analysis Results

The main outcomes of this meta-analysis for +936 C/T polymorphism are listed in Table 2. We found no evidence indicating that +936 C/T polymorphism were significantly associated with the risk of breast cancer in total population when using CC vs. TT (OR=1.01, 95% CI=0.96-1.06, $P_h=1.00$). To test whether there is a positive association in another contrast models we subsequently assumed the CC+CT vs. TT model, failing to find any significant association (OR=1.00, 95% CI=0.96-1.05, $P_h=1.0$). In terms of the remaining models tested, the statistical data showed no clear associations with the genetic risk of breast cancer (OR CC vs. CT+TT=1.02, 95% CI=0.98-1.07, $P_h=0.94$; OR C allele vs. T allele =1.01, 95% CI=0.98-1.04, $P_h=0.99$; OR CT vs. TT =1.01, 95% CI=0.93-1.09, $P_h=1.00$) (see Figure 2 and Figure 3).

Since the general analysis, which included data from all published studies, only showed insignificant results, we thus performed a stratified analysis by ethnicity to estimate the risk for each subgroup. However, the negative results were not altered in the analysis of Caucasian populations: OR CC vs. TT=1.00, 95% CI=0.95-1.05, $P_h=1.0$; OR CC+CT vs. TT=1.00, 95% CI=0.95-1.05, $P_h=1.0$; OR CC vs. CT+TT=1.02, 95% CI=0.97-1.07, $P_h=0.87$; OR C allele vs. T allele =1.01, 95% CI=0.98-1.04, $P_h=0.99$; OR CT vs. TT =0.99, 95% CI=0.90-1.08, $P_h=1.0$, and Asian populations OR CC vs. TT=1.00, 95% CI=0.95-1.05, $P_h=1.0$.

Table 1: Main characteristics of all studies included in the meta-analysis

| First author | Year | Country | Ethnicity | Control source | Genotyping method | HWE | Quality |
|--------------|------|---------|-----------|----------------|-------------------|-----|---------|
| Krippil35    | 2003 | Austria | Caucasian | Population    | PCR-RFLP          | 0.432 | High    |
| Jin16        | 2005 | Sweden  | Caucasian | Population    | PCR-RFLP          | 0.940 | High    |
| Kataoka17    | 2006 | China   | Asian     | Population    | DS                | 0.129 | High    |
| Jacobs16     | 2006 | USA     | Caucasian | Population    | TaqMan            | 0.602 | High    |
| Balasubramanian19 | 2007 | UK      | Caucasian | Population    | PCR-RFLP          | 0.842 | High    |
| Pharoah17    | 2007 | UK      | Caucasian | Population    | TaqMan            | 0.688 | High    |
| Eroglu20     | 2008 | Turkey  | Caucasian | Population    | PCR-RFLP          | 0.843 | High    |
| Jakubowska22 | 2008 | Poland  | Caucasian | Population    | PCR-RFLP          | 0.738 | High    |
| Lin23        | 2009 | China   | Caucasian | Hospital      | PCR-RFLP          | 0.023 | Low     |
| Jakubowska24 | 2009 | Poland  | Caucasian | Population    | PCR-RFLP          | 0.830 | High    |
| Oliveira25   | 2011 | Brazil  | Caucasian | Population    | PCR-RFLP          | 0.201 | High    |
| Luo26        | 2013 | China   | Asian     | Hospital      | PCR-RFLP          | 0.004 | Low     |

PCR: Polymerase chain reaction; PCR-RFLP: PCR-restriction fragment length polymorphism; DS: Direct sequencing; TaqMan: TaqManSNP; HWE: Hardy-Weinberg equilibrium


| Subgroups                  | OR (95% CI)               | P<sub>h</sub> | OR (95% CI)               | P<sub>h</sub> | OR (95% CI)               | P<sub>h</sub> | OR (95% CI)               | P<sub>h</sub> | OR (95% CI)               | P<sub>h</sub> |
|----------------------------|---------------------------|--------------|---------------------------|--------------|---------------------------|--------------|---------------------------|--------------|---------------------------|--------------|
| **Ethnicity**              |                           |              |                           |              |                           |              |                           |              |                           |              |
| Caucasian                  | 1.00 (0.95, 1.05)          | 1.00         | 1.00 (0.95, 1.05)          | 1.00         | 1.02 (0.97, 1.07)          | 0.87         | 1.01 (0.98, 1.04)          | 0.99         | 0.99 (0.90, 1.08)          | 1.00         |
| Asian                      | 1.04 (0.93, 1.16)          | 0.75         | 1.02 (0.93, 1.12)          | 0.79         | 1.02 (0.92, 1.13)          | 0.74         | 1.02 (0.96, 1.10)          | 0.68         | 1.07 (0.91, 1.26)          | 0.70         |
| **Control source**         |                           |              |                           |              |                           |              |                           |              |                           |              |
| Population                 | 1.00 (0.95, 1.05)          | 1.00         | 1.00 (0.96, 1.05)          | 1.00         | 1.02 (0.97, 1.07)          | 0.90         | 1.01 (0.98, 1.04)          | 0.99         | 1.00 (0.92, 1.09)          | 1.00         |
| Hospital                   | 1.04 (0.89, 1.21)          | 0.69         | 1.03 (0.90, 1.17)          | 0.73         | 1.07 (0.92, 1.23)          | 0.69         | 1.04 (0.95, 1.15)          | 0.97         | 1.07 (0.86, 1.34)          | 0.54         |
| **Quality**                |                           |              |                           |              |                           |              |                           |              |                           |              |
| High-quality               | 1.00 (0.95, 1.05)          | 1.00         | 1.00 (0.96, 1.05)          | 1.00         | 1.02 (0.97, 1.07)          | 0.90         | 1.01 (0.98, 1.04)          | 0.99         | 1.00 (0.92, 1.09)          | 1.00         |
| Low-quality                | 1.04 (0.89, 1.21)          | 0.69         | 1.03 (0.90, 1.17)          | 0.73         | 1.07 (0.92, 1.23)          | 0.69         | 1.04 (0.95, 1.15)          | 0.97         | 1.07 (0.86, 1.34)          | 0.54         |
| All                        | 1.01 (0.96, 1.06)          | 1.00         | 1.00 (0.96, 1.05)          | 1.00         | 1.02 (0.98, 1.07)          | 0.94         | 1.01 (0.98, 1.04)          | 0.99         | 1.01 (0.93, 1.09)          | 1.00         |

P<sub>h</sub>: P value of heterogeneity test; CI: Confidence interval; OR: Odds ratio

Similarly, the subgroups of hospital-based studies and population-based studies did not indicate any statistical evidence of significant breast cancer risk in relation to +936 C/T polymorphism at either the allelic or the genotypic level, nor did the final
analysis according to the quality of study reveal a statistically significant association (see table 2).

**Sensitivity Analysis and Potential Bias**

Notable alternation in combined effect estimates is a common event when performing meta-analysis, where all genetic association studies addressing the same topic were incorporated. To check whether the pooled ORs were significantly affected by some specific study, we carried out a sensitivity analysis through sequential exclusion of each independent study. The data remained stable during the analysis and it is the unaltered ORs that ensure the reliability of our findings (data not shown). In the Begg’s test, no evidence of publication bias was detected by visual inspection of the funnel plots. P values of Egger’s test also indicated that there was no significant publication bias across the included studies (t=-0.70, P=0.49 for CC vs. CT+TT) (see figure 4).

**Discussion**

Angiogenesis is a complicated process that could result in the formation of new blood vessels from pre-existing vasculature and thereby facilitates tumor growth, invasiveness, and development of metastasis. Data from experimental and clinical research have suggested that breast cancer is an angiogenesis-dependent disease closely associated with the serum level of VEGF. The potentially functional polymorphism in the VEGF gene, +936 C/T, was shown to correlate with lower VEGF production. It is the biological function of +936 C/T that leads to widespread attention towards the association between presence of +936 C/T polymorphism and breast cancer susceptibility. The published reports, however, have produced contentious and controversial results. Heterogeneous populations across the earlier association studies are a plausible explanation for the controversy. Other factors, such as varying study designs, number of individuals’ genotypes, different experimental methodologies may also contribute to the existing discrepancy among the studies.

Meta-analysis is an analytical tool that provides credible evidence for malignancy risk and well-characterized sequential variations by systematically summarizing existing data. In the current meta-analysis, we combined all available data on the association of +936 C/T
polymorphism with breast cancer risk. The pooled analysis showed that there was no increased or decreased risk of breast cancer in correlation with any genotype or allele at +936 C/T. The lack of association with breast cancer was also observed in subgroup analyses according to ethnicity. In subsequent stratification analyses by source of controls and quality of study, nor did we find a clear association, the same as the previous analyses.

Several meta-analyses have been carried out, in an attempt to identify the association of +936 C/T polymorphism with breast cancer risk. For instance, Gu et al., who analyzed 5,729 cases and 5,868 controls, found +936 C/T polymorphism may not contribute to breast cancer susceptibility. This study was later commented and the association was re-evaluated in a recent publication, in which several methodological issues were addressed. The re-evaluation nevertheless failed to find any significant association. Consistent with the two earlier assessments, several meta-analyses reported in 2011 did not lend support modified breast cancer risk in relation to +936 C/T genotypes.

One shared issue among these meta-analyses is that not all of the published data are included. In addition, there was substantial heterogeneity in each of the earlier meta-analyses, leading to an increased likelihood of biased results and hence making overall estimates less reliable. We enlarged our sample by incorporating all available data of the precious meta-analyses and the recently published studies, which helped to strengthen the statistical power and increased the precision of our findings, although we did not find any evidence of significant correlation. Another explanation may be that the polymorphism being investigated is a low-penetration genetic variant and the minor impact on breast cancer susceptibility cannot be detected unless a sufficiently large analysis is performed. It is therefore worthwhile to determine the genetic association in a large-scale study.

According to experimental evidence of functional genetic polymorphism at candidate locus, the VEGF +936T allele is speculated to have an association with predisposition to breast cancer. Activator protein 4 represents a helix-loop-helix transcription factor that upregulates serum levels of various cellular and viral genes through binding to particular sites, and it is these specific enhancer sites that modulate expression of VEGF. It has been shown that the +936T allele associated with reduced levels of VEGF abrogates activator protein 4 located in +936C allele, and thereby results in reduced VEGF expression. The +936T allele was also shown to have major effects on decreased uptake of 18F-fluorodeoxyglucose, a marker used to detect and diagnose breast cancer. It is therefore convincing that VEGF +936 C/T likely play a significant role in breast cancer susceptibility. If this speculation can be confirmed in the following studies, the +936 C/T polymorphism may be used as a biomarker for breast cancer.

We did not find a significant association.
between VEGF +936 C/T genotypes and breast cancer susceptibility in both Asian and Caucasian populations. The real association may be underestimated or even masked as a result of the small number of individuals in each ethnic group, the Asian group in particular (3,664 subjects). The minor allele frequency (+936 C/T) of Asian populations (18.2%) is higher as compared to the global frequency (14.9%, http://www.ncbi.nlm.nih.gov/snp). Available data show a connection between the +936T allele and lower VEGF protein levels that is related to increased susceptibility of breast cancer, suggesting VEGF +936 C/T polymorphism may have effect modification of breast cancer in Asians. Our findings are obviously inconsistent with this hypothesis of a positive association and the inconsistency hints that the relation merits further investigation.

Several limitations of this study should be considered in interpreting our results. Firstly, mild deviation from HWE that probably results from methodological issues such as inappropriate or erroneous genotyping, selection bias and population categorization was detected in two studies, and this may have influenced our results, despite no indicated apparent alterations in the data with or without them. Secondly, we carried out an exhaustive literature search in medicine-specific databases and supplemental hand-search, failing to identify the publications written in other languages except English. As a result, only published data in the English language were analyzed in this meta-analysis. Further, unpublished data and ongoing studies were not considered in the present work, which may lead to biased results. Third, crude OR instead of adjusted OR was calculated to assess the risk of breast cancer. Since there was no uniform standard definition for the independent studies and the calculation of ORs varies substantially across the published studies. For example, Luo et al. adjusted ORs for sex and age, whereas the adjustment factors in the study by Oliveira et al. were age and race. The wide difference makes it impossible to provide an OR based on adjusted factors. Finally, it is known that breast cancer is an etiologically heterogeneous malignancy. Genetic contribution alone only accounts a part for the breast pathogenesis and the exact molecular mechanism that underlies this invasive disease remain largely unknown. Available data documented a significant involvement of gene-to-gene and gene-to-environment interactions in the development and progression of breast cancer, which, however, cannot be confirmed in this analysis due to lack of the original data.

The aforementioned shortcomings, nevertheless, could not overshadow the strong points of this quantitative assessment. We performed a meta-analysis with a maximum sample to date and it is the sample sufficiency that has contributed to robust and convincing findings. Besides, we updated earlier meta-analyses with new information from subsequent published studies and determined that +936 C/T genotypes were not involved in breast cancer incidence. Last but not least, publication bias and inter-study inconsistency (heterogeneity) are two major problems when performing meta-analysis. The confounding factors appeared to have no notable influence on this investigation, as suggested in the analytical methods.

In summary, the accumulated evidence from prospective studies supports that +936 C/T polymorphism is not significantly associated with the risk of breast cancer. Well-designed studies with a larger number of samples and with gene-to-gene and gene-to-environment interactions are considered necessary to clarify the association.

This meta-analysis suggests that the common +936 C/T polymorphism may not be associated with breast cancer risk. This finding merits further research where the sample size is substantially large, and gene-gene and gene-environment interactions are considered.

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