Article

Receiver Operating Characteristic Prediction for Classification: Performances in Cross-Validation by Example

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Abstract: The stability of receiver operating characteristic in context of random split used in development and validation sets, as compared to the full models for three inflammatory ratios (neutrophil-to-lymphocyte (NLR), derived neutrophil-to-lymphocyte (dNLR) and platelet-to-lymphocyte (PLR) ratio) evaluated as predictors for metastasis in patients with colorectal cancer, was investigated. Data belonging to patients admitted with the diagnosis of colorectal cancer from January 2014 until September 2019 in a single hospital were used. There were 1688 patients eligible for the study, 418 in the metastatic stage. All investigated inflammatory ratios proved to be significant classification models on both the full models and on cross-validations (AUCs > 0.05). High variability of the cut-off values was observed in the unrestricted and restricted split (full models: 4.255 for NLR, 2.745 for dNLR and 255.56 for PLR; random splits: cut-off from 3.215 to 5.905 for NLR, from 2.625 to 3.575 for dNLR and from 134.67 to 335.9 for PLR), but with no effect on the models characteristics or performances. The investigated biomarkers proved limited value as predictors for metastasis (AUCs < 0.8), with largely sensitivity and specificity (from 33.3% to 79.2% for the full model and 29.1% to 82.7% in the restricted splits). Our results showed that a simple random split of observations, weighting or not the patients with and without metastasis, in a ROC analysis assures the performances similar to the full model, if at least 70% of the available population is included in the study.

Keywords: receiver operating characteristic; area under the curve; models performances; biomarkers; metastasis

1. Introduction

The receiver operating characteristic (ROC) has been introduced in medicine as a methodological tool used for the investigation of signal detection [1–3]. The method uses the probability of correct detection (as a binary outcome) against the probability of a false positive outcome. The ROC detection
method has been widely adopted in medicine and it is used to confirm (rule-in) or exclude (rule-out) the presence of a disease, as instruments of triage, monitorization, prognosis or screening [4]. The ROC analysis is particularly used in laboratory medicine to identify the threshold able to discriminate pathological versus normal values of biomarkers [5–8], images in radiology [9–12] or accounts in bioinformatics [13–15]. The methodology of ROC analysis is extensively described in the scientific literature, with emphasize on the need of a gold standard method in order to certify the presence/absence of the disease, to understand the effects of verification bias (application of the test only on those with known disease status [16,17]) or to verify indices and metrics of accuracy (also known as classification performances) [4,9,18–21]. A gold standard or reference test is defined as the best available method able to classify the subjects as presenting or not a certain disease [22]. Most of the time, the available reference method is not necessarily a perfect method of classification, so we have an imperfect standard [23] and the effects of the measurement errors have already been evaluated [24,25]. Latent class models (LCM), which combined multiple test results to estimate the accuracy of a test in the absence of a reference standard have also been proposed and their utility proved [26–28].

The concept of big data provides the input data for uni- and multivariate association towards precision medicine [29]. New methods and learning algorithms emerged for analysis, medical data interpretation [30] and the main challenges are validation and replication. The classification model is identified in a training or development set and validated on a validation set [31]. The cross-validation, repeatedly randomised division of a dataset into development and validation sets, is applied to test the efficiency of the model on validation of external data [31–33]. Several cross-validation approaches, such as leave-one-out (LOO—one observation is left out of the development set) [34], leave-p-out (p observations are left out of the development set) [35], k-fold (the sample is split in k equal size sub-samples and one sub-sample is being repeated k-times, the one used for validation) [36], bootstrap method [37] or holdout method (also known as random splitting—observations are randomly split into two sets, development and validation set) [38] are used.

Colorectal cancer is the third most common type of cancer in men and second in women. The death rate from colorectal cancer is 8.5%. Surgical resection remains the only curative option for these patients, although a number of them develop metastases in the first five years after surgery [39]. Tumour size, tumour location, venous invasion, degree of differentiation proved valuable abilities to correctly predict local recurrences, metastasis and survival rate of colorectal cancer patients [40]. New biomarkers, such as chemokine ligand 7 (CXCL7) serum concentration have been tested for diagnostic accuracy [41], but the costs of such new biomarkers are generally higher than the conventional ones and health insurance systems do not support them. Consequently, a low-cost and time-bound predictor would be desirable. Pine et al. showed that tumour progression and prognosis are bound to tumour characteristics, as well as to the inflammatory response [42]. The inflammatory response is reflected by the level of neutrophils, lymphocytes, platelets, albumin and C-reactive protein, their measurement being current practice. Neutrophil-to-lymphocyte ratio (NLR), platelet-to-lymphocyte ratio (PLR), lymphocyte-to-monocyte ratio (LMR) along with others, proved predictors for survival of patients with colorectal cancer [43–45]. Xia et al. reported NLR (cut-off = 2.8, AUC (Area Under the Curve) = 0.711, p = 0.007), LMR (cut-off = 3.9, AUC = 0.679, p = 0.023) and PNI (Prognostic Nutritional Index, cut-off = 47.1, AUC = 0.746, p = 0.002) as predictors for post-operative complications in patients with T1-T2 rectal cancer [46]. NLR (cut-off = 2.15, AUC = 0.790, 95% CI (0.736 to 0.884), CI = confidence interval) and PLR (cut-off = 123, AUC = 0.846, 95% CI (0.801 to 0.891)) presented performance as diagnostic markers for colorectal cancer as reported by Stojkovic Lalosevic [47]. Different cut-off values of the biomarkers are reported in the range from moderate to low performances [43–50]. In the light of the state-of-the-art, our goals were two-folded, first to evaluate the performances of ROC models using three ratios associated with metastasis on unrestricted and restricted random cross-validation and secondly, to demonstrate the impact of these two randomizations approaches on the performances of the classification models.
2. Materials and Methods

The study was conducted according to the principles of the Declaration of Helsinki. It was approved by the Ethical Committee of “Iuliu Hatieganu” University of Medicine and Pharmacy Cluj-Napoca (approval no. 492/21.11.2019) and the Ethical Committee of “Prof. Dr. Octavian Fodor” Regional Institute of Gastroenterology and Hepatology Cluj-Napoca (approval no. 17517/20.12.2019). All participants signed the written informed consent before enrolment in the study.

2.1. Dataset

The dataset was represented by a retrospective collection of absolute neutrophil, lymphocyte, leucocyte, platelet counts of patients admitted at Third Surgical Clinic, “Prof. Dr. Octavian Fodor” Regional Institute of Gastroenterology and Hepatology Cluj-Napoca with the diagnosis of colorectal cancer from January 2014 until September 2019. The presence of metastasis was the outcome variable in this study. The metastatic stage was documented by imaging explorations, namely computer tomography (CT), magnetic resonance imaging (MRI) or contrast enhanced ultrasound. Raw data belonging to the patients with histopathological diagnosis of colorectal cancer were evaluated. Records with incomplete data (e.g., laboratory results, TNM classification of the disease—T size or extent of the primary tumor, N degree of spread to regional lymph nodes, and M presence of distant metastasis) were excluded.

The absolute counts were the input data for calculation of the following ratios, used as predictors of metastasis in this study:

\[ \text{NLR} = \frac{\text{absolute neutrophil count}}{\text{absolute lymphocyte count}} \]  
\[ \text{dNLR} = \frac{\text{absolute neutrophil count}}{(\text{absolute leucocyte count}) - (\text{absolute neutrophil count})} \]  
\[ \text{PLR} = \frac{\text{absolute platelet count}}{\text{absolute lymphocyte count}} \]

where NLR = neutrophil-to-lymphocyte ratio, dNLR = derived neutrophil-to-lymphocyte ratio and PLR = platelet-to-lymphocyte ratio.

2.2. Methods

The receiver operating characteristic (ROC) analysis was conducted using three scenarios to test the classification abilities of the predictive models incorporating the studied ratios, which are presented in Equations (1)–(3) regarding the presence of metastasis (Figure 1):

- First scenario (\(1\)): the whole sample was used to generate the classification model, the control model.
- Second scenario (\(2\)): thirteen random sets were generated (1688 patients) by specifying the desired percentage of subjects and used to identify the classification models. No restrictions were imposed regarding the percentage of patients with and without metastasis and for this reason the generated sets were named unrestricted (\(UR_{x,x}\), where \(1 \leq x \leq 13\)). The following percentages were used: 70% for \(UR_{01}\), 65% for \(UR_{02}\), 60% for \(UR_{03}\), 55% for \(UR_{04}\), 50% for \(UR_{05}\), 45% for \(UR_{06}\), 40% for \(UR_{07}\), 35% for \(UR_{08}\), 30% for \(UR_{09}\), 25% for \(UR_{10}\), 20% for \(UR_{11}\), 15% for \(UR_{12}\), 10% for \(UR_{13}\).
- Third scenario (\(3\)): five sets, each with a development and validation group were randomly generated by weighting the percentage of patients with and without metastasis. These sets were name restricted (\(R\_\text{Runy}\), \(1 \leq y \leq 5\); with 70% in development set & 30% in validation set). The model was generated using the development set (Se, Sp, Acc—Accuracy) and tested on validation set (PPV—Positive Predictive Value and NPV—Negative Predictive Value).
The performances of the models were evaluated whenever the AUC proved statistically significant and the Gini index (GI) was higher than 0 (GI equal with 0 indicates a random model). Ten metrics (see Figure 1) were used to prove the accuracy of a classification model. The sensitivity and specificity, important metrics widely used to show models accuracy [4,5,11,14,51,52]. The F1-score \((2 \times ((Se \times PPV)/(Se + PPV)))\) was calculated by combining the performances in the development set with those in the validation set in restricted runs. Matthews correlation coefficient \(\text{MCC} = (TP \times TN - FP \times FN)/(TP + FN) \times (TN + FP) \times (TN + FN))\) confers a more suitable balance to all four confusion matrix categories (\(TP—true\ positive, \ TN—true\ negative, \ FP—false\ positive, \ FN—false\ negative\)) and it is considered a more specific score [53], compared to the model accuracy or F1-Score, which are overoptimistic estimators [54].

The ROC analysis was conducted with SPSS v. 26 (trial version) under the non-parametric assumptions. The cut-off value for each biomarker in each run was used to create a derivate dichotomial variable (presence of metastasis for values equal to or higher than the cut-off). Youden’s index \((J = \text{max})\), which maximizes the distance from the diagonal (random classification model) [55] was used to identify the cut-off values. The tangent method \((d = \sqrt{(1 - \text{Se})^2 + (1 - \text{Sp})^2})\), \(d = \text{min}\) [56] was also used to identify the cut-off values in the control model. The observed confusion matrix was generated and the true positive (TP), false positive (FP), false negative (FN) and true negative (TN) values were used to calculate the performances of the models using the following online resource: https://statpages.info/ctab2x2.html (accessed 3 August 2020). Clinical utility index (CUI) was calculated for the control model using the online resource available at http://www.clinicalutility.co.uk/ (accessed 30 September 2020). A fair utility for case-finding (+CUI) or screening (-CUI) is seen if \(0.49 \leq \text{CUI} < 0.69\), a good utility if \(0.69 \leq \text{CUI} < 0.81\) and an excellent utility if \(\text{CUI} \geq 0.81\) [57].

3. Results

3.1. First Scenario: Full Predictive Models

Significant contribution in identifying the presence of metastasis (AUC significantly different by 0.5) was identified for all investigated ratios, when the whole sample was used (Table 1).
Table 1. Characteristics of the full (control) models.

|              | AUC [95%CI] | StdErr | p-Value | Gini Index | Cut-Off Value * |
|--------------|-------------|--------|---------|------------|-----------------|
| NLR          | 0.647 [0.616 to 0.679] | 0.0161 | <0.0001 | 0.295      | 4.255           |
| dNLR         | 0.634 [0.602 to 0.666] | 0.0162 | <0.0001 | 0.268      | 2.745           |
| PLR          | 0.568 [0.535 to 0.601] | 0.0168 | <0.0001 | 0.136      | 255,560         |

NLR = neutrophil-to-lymphocyte ratio; dNLR = derived neutrophil-to-leucocyte ratio; PLR = platelet-to-lymphocyte ratio; AUC = area under the curve; [95%CI] = lower bound to upper bound of the AUC 95% confidence interval; StdErr = standard error; * J = max (Youden index), where Youden index = (Se + Sp − 1); Se = sensitivity; Sp = specificity; d = min.

The performances metrics of the control models showed modest or low classification abilities of the investigated ratios (Table 2), NLR being the most performant biomarker for metastasis.

Table 2. Performances of the (control) models (reported in Table 1).

|              | NLR          | dNLR         | PLR          |
|--------------|--------------|--------------|--------------|
|              | cut-off threshold by J = max |              |              |
| TP           | 220          | 228          | 139          |
| TN           | 902          | 860          | 1006         |
| FP           | 368          | 410          | 264          |
| FN           | 198          | 190          | 279          |
| Se           | 52.6 [47.8 to 57.4] | 54.5 [49.8 to 59.3] | 33.3 [28.7 to 37.8] |
| Sp           | 71.0 [68.5 to 73.5] | 67.7 [65.1 to 70.3] | 79.2 [77.0 to 81.4] |
| Acc          | 66.5 [64.2 to 68.7] | 64.5 [52.2 to 66.7] | 67.8 [65.6 to 70.1] |
| PPV          | 37.4 [34.4 to 40.3] | 35.7 [33.0 to 38.5] | 24.8 [22.8 to 26.9] |
| NPV          | 82 [80.4 to 83.6] | 81.9 [80.2 to 83.6] | 151 [11.19 to 18.9] |
| +LR          | 1.8 [1.6 to 2.1] | 1.7 [1.5 to 1.9] | 1.6 [1.3 to 1.9] |
| -LR          | 0.7 [0.6 to 0.7] | 0.7 [0.6 to 0.7] | 0.8 [0.8 to 0.9] |
| MCR          | 33.5 [31.5 to 35.6] | 35.5 [33.5 to 37.6] | 32.2 [30.2 to 34.1] |
| DOR          | 2.7 [2.2 to 3.4] | 2.5 [2 to 3.2] | 1.9 [1.5 to 2.4] |
| NND          | 4.2 [3.4 to 5.5] | 4.5 [3.6 to 6] | 8 [5.7 to 13.5] |
| +CUI         | 0.197 [0.144 to 0.250] | 0.195 [0.413 to 0.247] | 0.115 [0.059 to 0.171] |
| -CUI         | 0.582 [0.561 to 0.604] | 0.555 [0.532 to 0.577] | 0.620 [0.601 to 0.644] |
| F1-Score     | 51.4          | 49.7          | 18.8          |
| MCC          | 0.214         | 0.198         | 0.126         |

|              | cut-off threshold by d = min |              |              |
| TP           | 244          | 235          | 229          |
| TN           | 810          | 836          | 691          |
| FP           | 460          | 434          | 579          |
| FN           | 174          | 183          | 189          |
| Se           | 58.4 [54.1 to 62.5] | 56.2 [52.0 to 60.4] | 54.8 [50.5 to 59] |
| Sp           | 63.8 [62.4 to 65.1] | 65.8 [64.4 to 67.2] | 54.4 [53 to 55.8] |
| Acc          | 62.6 [60.3 to 64.5] | 63.4 [61.4 to 65.5] | 54.5 [52.4 to 56.6] |
| PPV          | 34.7 [32.1 to 37.1] | 35.1 [32.5 to 37.6] | 28.3 [26.1 to 30.5] |
| NPV          | 82.3 [80.5 to 84.1] | 82 [80.3 to 83.7] | 78.5 [76.5 to 80.5] |
| +LR          | 1.6 [1.4 to 1.8] | 1.6 [1.5 to 1.8] | 1.2 [1.1 to 1.3] |
| -LR          | 0.7 [0.6 to 0.7] | 0.7 [0.6 to 0.7] | 0.8 [0.7 to 0.9] |
| MCR          | 37.6 [35.5 to 39.7] | 36.6 [34.5 to 38.6] | 45.5 [43.4 to 47.6] |
| DOR          | 2.469 [1.958 to 3.115] | 2.474 [1.962 to 3.119] | 1.4 [1.2 to 1.8] |
| NND          | 4.51 [3.62 to 6.05] | 4.54 [3.63 to 6.09] | 10.9 [6.8 to 28.5] |
| +CUI         | 0.202 [0.152 to 0.253] | 0.197 [0.416 to 0.249] | 0.155 [0.107 to 0.203] |
| -CUI         | 0.525 [0.502 to 0.548] | 0.540 [0.517 to 0.563] | 0.427 [0.402 to 0.453] |
| F1-Score     | 43.53         | 43.21         | 37.3          |
| MCC          | 0.194         | 0.195         | 0.079         |

NLR = neutrophil-to-lymphocyte ratio; dNLR = derived neutrophil-to-leucocyte ratio; PLR = platelet-to-lymphocyte ratio; TP = true positive; TN = true negative; FP = false positive; FN = false negative; Se = sensitivity; Sp = specificity; Acc = accuracy; LR = likelihood ratio; MCR = miss-classification rate; DOR = diagnostic odds ratio; NND = number needed to diagnose.
3.2. Second Scenario: Unrestricted Random Samples

At least one unrestricted random sample retrieves the same Youden’s cut-off value for each investigated ratio as the full predicted model (URun01 for NLR and dNLR; URun02 for dNLR; URun05 for PLR; see Tables 1 and 3).

Table 3. Number of total subjects in the run, distribution among those with and without metastasis, cut-off values and models characteristics for unrestricted random samples.

| Run  | N [%] with Metastasis | N [%] without Metastasis | NLR | dNLR | PlR |
|------|-----------------------|--------------------------|-----|------|-----|
| URun01—1222 patients (72.4%); 307 with (73.4%) and 915 without (72.0%) metastasis | NLR = 0.658 [0.621 to 0.695] 0.019 <0.0001 0.316 4.255 | dNLR = 0.648 [0.611 to 0.685] 0.019 <0.0001 0.297 2.745 | PlR = 0.565 [0.527 to 0.604] 0.020 0.0009 0.131 246.650 |
| URun02—1050 patients (62.2%); 267 with (63.9%) and 783 without (61.7%) metastasis | NLR = 0.643 [0.602 to 0.683] 0.0208 <0.0001 0.285 3.635 | dNLR = 0.625 [0.584 to 0.666] 0.0208 <0.0001 0.250 2.745 | PlR = 0.564 [0.521 to 0.607] 0.0211 <0.0001 0.171 257.365 |
| URun03—1003 patients (59.4%); 258 with (61.7%) and 745 without (58.7%) metastasis | NLR = 0.649 [0.609 to 0.689] 0.0203 <0.0001 0.298 3.625 | dNLR = 0.635 [0.595 to 0.676] 0.0205 <0.0001 0.271 2.755 | PlR = 0.584 [0.543 to 0.626] 0.0212 <0.0001 0.169 246.410 |
| URun04—905 patients (53.6%); 218 with (52.2%) and 687 without (54.1%) metastasis | NLR = 0.678 [0.637 to 0.719] 0.0210 <0.0001 0.356 3.625 | dNLR = 0.656 [0.615 to 0.698] 0.0214 <0.0001 0.313 2.655 | PlR = 0.576 [0.531 to 0.621] 0.0229 <0.0001 0.152 185.145 |
| URun05—848 patients (50.2%); 207 with (49.5%) and 641 without (50.5%) metastasis | NLR = 0.623 [0.577 to 0.670] 0.0235 <0.0001 0.247 4.465 | dNLR = 0.597 [0.551 to 0.644] 0.0237 <0.0001 0.195 2.985 | PlR = 0.558 [0.512 to 0.605] 0.0236 0.0135 0.117 255.680 |
| URun06—772 patients (45.7%); 201 with (48.1%) and 571 without (45.0%) metastasis | NLR = 0.639 [0.595 to 0.683] 0.0226 <0.0001 0.278 3.635 | dNLR = 0.621 [0.576 to 0.666] 0.0229 <0.0001 0.242 2.670 | PlR = 0.576 [0.53 to 0.622] 0.0236 0.0013 0.152 162.995 |
| URun07—664 patients (39.3%); 154 with (36.8%) and 608 without (52.3%) metastasis | NLR = 0.654 [0.602 to 0.705] 0.0264 <0.0001 0.307 5.905 | dNLR = 0.637 [0.585 to 0.689] 0.0266 <0.0001 0.274 3.135 | PlR = 0.574 [0.520 to 0.627] 0.0272 0.0069 0.147 255.845 |
| URun08—597 patients (35.4%); 158 with (37.8%) and 439 without (34.6%) metastasis | NLR = 0.616 [0.564 to 0.668] 0.0266 <0.0001 0.233 5.855 | dNLR = 0.607 [0.554 to 0.66] 0.0270 <0.0001 0.214 2.740 | PlR = 0.53 [0.476 to 0.584] 0.0276 0.2726 0.061 335.900 |
| URun09—507 patients (30.0%); 120 with (28.7%) and 387 without (30.5%) metastasis | NLR = 0.639 [0.586 to 0.698] 0.0299 <0.0001 0.276 3.645 | dNLR = 0.622 [0.562 to 0.682] 0.0304 <0.0001 0.244 3.575 | PlR = 0.56 [0.502 to 0.619] 0.0299 0.0436 0.121 150.380 |
| URun10—403 patients (23.9%); 107 with (25.6%) and 296 without (23.3%) metastasis | NLR = 0.684 [0.623 to 0.746] 0.0316 <0.0001 0.369 4.545 | dNLR = 0.673 [0.612 to 0.734] 0.0311 <0.0001 0.346 2.645 | PlR = 0.608 [0.541 to 0.675] 0.0343 0.0016 0.216 265.875 |
| URun11—320 patients (19.0%); 78 with (18.7%) and 242 without (19.1%) metastasis | NLR = 0.665 [0.591 to 0.74] 0.0359 <0.0001 0.331 3.635 | dNLR = 0.661 [0.594 to 0.738] 0.0381 <0.0001 0.308 3.575 | PlR = 0.574 [0.501 to 0.647] 0.0372 0.0480 0.147 258.540 |
| URun12—260 patients (15.4%); 67 with (16.0%) and 193 without (15.2%) metastasis | NLR = 0.627 [0.547 to 0.708] 0.0410 0.0019 0.255 4.250 | dNLR = 0.619 [0.538 to 0.7] 0.0413 0.0039 0.238 3.145 | PlR = 0.575 [0.489 to 0.661] 0.0441 0.0885 0.150 265.230 |
| URun13—136 patients (8.1%); 34 with (8.1%) and 102 without (8.0%) metastasis | NLR = 0.674 [0.568 to 0.780] 0.0541 0.0013 0.348 3.215 | dNLR = 0.666 [0.536 to 0.775] 0.0561 0.0032 0.331 3.080 | PlR = 0.521 [0.408 to 0.634] 0.0576 0.7193 0.042 134.670 |

n = 1688; AUC = area under the curve; [95%CI] = lower bound to upper bound of the AUC 95% confidence interval; StdErr = standard error; * Youden index (Se + Sp - 1), where Se = sensitivity; Sp = specificity; NLR = neutrophil-to-lymphocyte ratio; dNLR = derived neutrophil-to-leucocyte ratio; PlR = platelet-to-lymphocyte ratio.
The obtained cut-off values were lower than the values reported in Table 1 for NRL (8/13, 61.5%; Table 3). Changes in the significance of the AUC were observed only for PLR for unrestricted random samples with 16% of the subjects in the development set (Table 3).

### 3.3. Third Scenario: Restricted Random Samples

The classification models in cross-validation, with 70% of the available data in the development set and 30% in the validation set, showed similar performances of AUCs for NLR and dNLR ($p$-values in the same range), but with AUCs close to 0.5 for PLR (in all RRun, Table 4).

**Table 4.** Models characteristics in restricted random samples.

|               | AUC [95%CI] | StdErr | $p$-Value | Gini Index | Cut-Off Value * |
|---------------|-------------|--------|-----------|------------|-----------------|
| **RRun01**    |             |        |           |            |                 |
| NLR           | 0.647 [0.610 to 0.684] | 0.0189 | <0.0001   | 0.295      | 4.225           |
| dNLR          | 0.629 [0.591 to 0.666] | 0.0190 | <0.0001   | 0.257      | 2.755           |
| PLR           | 0.559 [0.520 to 0.598] | 0.0199 | 0.0032    | 0.117      | 268.675         |
| **RRun02**    |             |        |           |            |                 |
| NLR           | 0.638 [0.600 to 0.676] | 0.0195 | <0.0001   | 0.275      | 4.065           |
| dNLR          | 0.616 [0.577 to 0.654] | 0.0195 | <0.0001   | 0.231      | 2.835           |
| PLR           | 0.564 [0.525 to 0.603] | 0.0200 | 0.0013    | 0.128      | 225.480         |
| **RRun03**    |             |        |           |            |                 |
| NLR           | 0.650 [0.612 to 0.689] | 0.0195 | <0.0001   | 0.300      | 4.475           |
| dNLR          | 0.635 [0.596 to 0.673] | 0.0196 | <0.0001   | 0.269      | 3.135           |
| PLR           | 0.575 [0.536 to 0.614] | 0.0200 | 0.0002    | 0.151      | 255.560         |
| **RRun04**    |             |        |           |            |                 |
| NLR           | 0.643 [0.606 to 0.681] | 0.0191 | <0.0001   | 0.287      | 3.635           |
| dNLR          | 0.626 [0.589 to 0.664] | 0.0193 | <0.0001   | 0.253      | 3.075           |
| PLR           | 0.570 [0.531 to 0.609] | 0.0200 | 0.0005    | 0.140      | 255.845         |
| **RRun05**    |             |        |           |            |                 |
| NLR           | 0.634 [0.596 to 0.672] | 0.0192 | <0.0001   | 0.268      | 4.225           |
| dNLR          | 0.619 [0.581 to 0.657] | 0.0193 | <0.0001   | 0.238      | 2.755           |
| PLR           | 0.570 [0.532 to 0.609] | 0.0196 | 0.0003    | 0.141      | 261.695         |

NLR = neutrophil-to-lymphocyte ratio; dNLR = derived neutrophil-to-leucocyte ratio; PLR = platelet-to-lymphocyte ratio; AUC = area under the curve; [95%CI] = lower bound to upper bound of the AUC 95% confidence interval; StdErr = standard error; * Youden index (Se + Sp – 1), where Se = sensitivity; Sp = specificity.

The true positive, true negative, false positive and false negative in restricted random samples varied in the same way in the development and validation sets (Table 5). The above-mentioned values varied between runs in different amounts, ranging, for example, in the case of true positives from a difference between maximum and minimum form of 3.4% observed for PLR and 46.2% observed for AGR (Table 5).

The cross-validation ROC analysis applied on restricted random samples identified the NLR and dNLR as the most potent markers for metastasis according to the highest F1-Score (Table 6). However, better performances are obtained when the whole sample was investigated for NLR and dNLR, while on the opposite, better performances are obtained in cross-validation for PLR, all of them showing the instability of the classification model dictated by the input data (Tables 2 and 6).
Table 5. Number and frequencies of true positive, true negative, false positive and false negative in the restricted random samples according to run.

| Ratio | Run   | Development | Validation |
|-------|-------|-------------|------------|
|       |       | TP          | TN          | FP          | FN          | TP          | TN          | FP          | FN          |
| NLR   | RRun01| 153 (52.4)  | 629 (70.8)  | 259 (29.2)  | 139 (47.6)  | 69 (54.8)   | 264 (69.1)  | 118 (30.9)  | 57 (45.2)   |
|       | RRun02| 158 (54.1)  | 616 (68.6)  | 282 (31.4)  | 134 (45.9)  | 72 (57.1)   | 246 (66.1)  | 126 (33.9)  | 54 (42.9)   |
|       | RRun03| 145 (49.7)  | 669 (75.3)  | 219 (24.7)  | 147 (50.3)  | 61 (48.4)   | 275 (72.0)  | 107 (28.0)  | 65 (51.6)   |
|       | RRun04| 182 (62.3)  | 534 (60.1)  | 354 (39.9)  | 110 (37.7)  | 81 (64.3)   | 226 (59.2)  | 156 (40.8)  | 45 (35.7)   |
|       | RRun05| 150 (51.4)  | 618 (69.6)  | 270 (30.4)  | 142 (48.6)  | 72 (57.1)   | 275 (72.0)  | 107 (28.0)  | 54 (42.9)   |
| dNLR  | RRun01| 153 (52.4)  | 606 (68.2)  | 282 (31.8)  | 139 (47.6)  | 74 (58.7)   | 256 (67)    | 126 (33)    | 52 (41.3)   |
|       | RRun02| 147 (50.3)  | 630 (70.2)  | 268 (29.8)  | 145 (49.7)  | 70 (55.6)   | 261 (70.2)  | 111 (29.8)  | 56 (44.4)   |
|       | RRun03| 127 (43.5)  | 703 (79.2)  | 185 (20.8)  | 165 (56.5)  | 53 (42.1)   | 289 (75.7)  | 93 (24.3)   | 73 (57.9)   |
|       | RRun04| 124 (42.5)  | 692 (77.9)  | 196 (22.1)  | 168 (57.5)  | 81 (64.3)   | 226 (59.2)  | 156 (40.8)  | 45 (35.7)   |
|       | RRun05| 152 (52.1)  | 597 (67.2)  | 291 (32.8)  | 140 (47.9)  | 75 (59.5)   | 265 (69.4)  | 117 (30.6)  | 51 (40.5)   |

NLR = neutrophil-to-lymphocyte ratio; dNLR = derived neutrophil-to-leucocyte ratio; PLR = platelet-to-lymphocyte ratio; TP = true positive; TN = true negative; FP = false positive; FN = false negative.

Table 6. Models performances in restricted random samples according to run.

| NLR | dNLR | PLR | Run   |
|-----|------|-----|-------|
|     |      |     | RRun01|
| Se  | 52.4 [47.4 to 57.4] | 29.5 [25.1 to 34] | 52.4 [47.3 to 57.4] |
| Sp  | 70.8 [69.2 to 72.5] | 82.1 [80.7 to 83.6] | 68.2 [66.6 to 69.9] |
| Acc | 66.3 [63.8 to 68.7] | 69.1 [66.9 to 71.3] | 64.3 [61.8 to 66.8] |
| +LR | 1.8 [1.54 to 2] | 1.65 [1.42 to 2] | 1.65 [1.42 to 2] |
| -LR | 0.67 [0.59 to 0.76] | 0.86 [0.79 to 0.93] | 0.7 [0.61 to 0.79] |
| MCR | 33.7 [31.3 to 36.2] | 30.9 [28.7 to 33.1] | 35.7 [33.2 to 38.2] |
| DOR | 2.7 [2 to 3.5] | 1.9 [1.4 to 2.6] | 2.4 [1.8 to 3.1] |
| NND | 3 [2.8 to 3.2] | 3.2 [3 to 3.5] | 2.8 [2.6 to 3] |
| MCC_dev | 0.025 | 0.016 | 0.022 |
| PPV | 36.9 [31.6 to 42] | 39.6 [30.9 to 48.6] | 37 [32 to 41.7] |
| NPV | 82.2 [79.2 to 85.2] | 78.6 [76.6 to 80.7] | 83.1 [79.9 to 86.2] |
| MCC_val | 0.038 | 0.032 | 0.041 |
| F1-Score | 43.3 | 33.8 | 43.4 |

| NLR | dNLR | PLR | Run02 |
|-----|------|-----|-------|
| Se  | 54.1 [49 to 59.1] | 37.0 [32.3 to 41.8] | 50.3 [45.3 to 55.3] |
| Sp  | 68.6 [66.9 to 70.2] | 76.7 [75.1 to 78.3] | 70.1 [68.5 to 71.7] |
| Acc | 65 [62.6 to 67.5] | 66.7 [64.4 to 69.1] | 65.3 [62.8 to 67.7] |
| +LR | 1.72 [1.48 to 2] | 1.59 [1.3 to 2] | 1.69 [1.44 to 2] |
| -LR | 0.67 [0.58 to 0.76] | 0.82 [0.74 to 0.9] | 0.71 [0.62 to 0.8] |
| MCR | 35 [32.5 to 37.4] | 33.3 [30.9 to 35.6] | 34.7 [32.3 to 37.2] |
| DOR | 2.6 [1.9 to 3.4] | 3 [2.8 to 3.2] | 2.4 [1.8 to 3.1] |
| NND | 2.9 [2.7 to 3.1] | 0.013 | 2.9 [2.7 to 3.1] |
| MCC_dev | 0.024 | 0.013 | 0.022 |
| PPV | 36.4 [31.4 to 41.2] | 34.8 [28.1 to 41.6] | 38.7 [33.2 to 43.9] |
| NPV | 82 [78.7 to 85.2] | 78.3 [75.8 to 81] | 82.3 [79.2 to 85.3] |
| MCC_val | 0.037 | 0.025 | 0.042 |
| F1-Score | 43.5 | 35.9 | 43.7 |
Table 6. Cont.

|               | NLR       | dNLR      | PLR       |
|---------------|-----------|-----------|-----------|
| **RRun03**    |           |           |           |
| Se            | 49.7 [44.7 to 54.6] | 43.5 [38.7 to 48.3] | 32.5 [28 to 37.2] |
| Sp            | 75.3 [73.7 to 76.9] | 79.2 [77.6 to 80.7] | 81.6 [80.2 to 83.2] |
| Acc           | 69 [66.5 to 71.4] | 70.3 [68 to 72.7] | 69.5 [67.3 to 71.8] |
| +LR           | 2.01 [1.7 to 2] | 2.09 [1.73 to 3] | 1.77 [1.41 to 2] |
| -LR           | 0.67 [0.59 to 0.75] | 0.71 [0.64 to 0.79] | 0.83 [0.76 to 0.9] |
| MCR           | 31 [28.6 to 33.5] | 29.7 [27.3 to 32] | 30.5 [28.2 to 32.7] |
| DOR           | 3 [2.3 to 4] | 2.9 [2.2 to 3.9] | 2.1 [1.6 to 2.9] |
| NND           | 3.2 [3 to 3.5] | 3.4 [3.1 to 3.7] | 3.3 [3.1 to 3.5] |
| MCC_{dev}     | 0.027 | 0.026 | 0.018 |
| PPV           | 36.3 [30.5 to 42] | 36.3 [29.9 to 42.8] | 32.5 [25.3 to 40.2] |
| NPV           | 80.9 [78 to 83.7] | 79.8 [77.2 to 82.5] | 77.6 [75.4 to 79.9] |
| MCC_{val}     | 0.033 | 0.031 | 0.019 |
| F1-Score      | 42 | 39.6 | 32.5 |
| **RRun04**    |           |           |           |
| Se            | 62.3 [57.2 to 67.2] | 42.5 [37.6 to 47.3] | 32.2 [27.7 to 36.8] |
| Sp            | 60.1 [58.5 to 61.7] | 77.9 [76.3 to 79.5] | 81.3 [79.8 to 82.8] |
| Acc           | 60.7 [58.2 to 63.1] | 69.2 [66.8 to 71.5] | 69.2 [66.9 to 71.4] |
| +LR           | 1.56 [1.38 to 2] | 1.92 [1.59 to 2] | 1.72 [1.37 to 2] |
| -LR           | 0.63 [0.53 to 0.73] | 0.74 [0.66 to 0.82] | 0.83 [0.76 to 0.91] |
| MCR           | 39.3 [36.9 to 41.8] | 30.8 [28.5 to 33.2] | 30.8 [28.6 to 33.1] |
| DOR           | 2.5 [1.9 to 3.3] | 2.6 [1.9 to 3.5] | 2.1 [1.5 to 2.8] |
| NND           | 2.5 [2.4 to 2.7] | 3.2 [3 to 3.5] | 3.2 [3 to 3.5] |
| MCC_{dev}     | 0.024 | 0.023 | 0.018 |
| PPV           | 34.2 [30 to 38.1] | 38 [31.9 to 44] | 33.6 [26.2 to 41.5] |
| NPV           | 83.4 [79.7 to 86.8] | 81.1 [78.4 to 83.9] | 77.8 [75.6 to 80.1] |
| MCC_{val}     | 0.038 | 0.037 | 0.021 |
| F1-Score      | 44.2 | 40.1 | 32.9 |
| **RRun05**    |           |           |           |
| Se            | 51.4 [46.3 to 56.4] | 52.1 [47 to 57.1] | 29.1 [24.8 to 33.6] |
| Sp            | 69.6 [67.9 to 71.2] | 67.2 [65.6 to 68.9] | 82.7 [81.2 to 84.1] |
| Acc           | 65.1 [62.6 to 67.6] | 63.5 [61 to 66] | 69.4 [67.3 to 71.6] |
| +LR           | 1.69 [1.45 to 2] | 1.59 [1.36 to 2] | 1.68 [1.32 to 2] |
| -LR           | 0.7 [0.61 to 0.79] | 0.71 [0.62 to 0.81] | 0.86 [0.79 to 0.93] |
| MCR           | 34.9 [32.4 to 37.4] | 36.5 [34 to 39] | 30.6 [28.4 to 32.7] |
| DOR           | 2.4 [1.8 to 3.2] | 2.2 [1.7 to 2.9] | 2.1 [1.4 to 2.7] |
| NND           | 2.9 [2.7 to 3.1] | 2.7 [2.6 to 2.9] | 3.3 [3.1 to 3.5] |
| MCC_{dev}     | 0.022 | 0.016 | 0.020 |
| PPV           | 40.2 [34.7 to 45.5] | 39.1 [33.9 to 43.9] | 36.1 [28.7 to 43.9] |
| NPV           | 83.6 [80.6 to 86.4] | 83.9 [80.7 to 86.8] | 78.7 [76.4 to 81] |
| MCC_{val}     | 0.047 | 0.047 | 0.027 |
| F1-Score      | 45.1 | 44.7 | 32.2 |

NLR = neutrophil-to-lymphocyte ratio; dNLR = derived neutrophil-to-leucocyte ratio; PLR = platelet-to-lymphocyte ratio; Se = sensitivity; Sp = specificity; Acc = accuracy; +LR = positive likelihood ratio; -LR = negative likelihood ratio; PPV = positive predictive value; NPV = negative predictive value; MCR = miss-classification rate; DOR = diagnostic odds ratio; NND = number needed to diagnose/predict; MCC = Matthews correlation coefficient; dev = development set; val = validation set.

The visual representation of the AUCs for each scenario related to the full model for NLR, dNLR, and PLR showed higher variability in the case of unrestricted random samples (Figure 2). The AUC representation of each run as compared to the full model is available in the Supplementary Material (Figures S1–S3).
Figure 2. The pattern of the AUCs: (a) neutrophil-to-lymphocyte ratio (NLR), (b) derived neutrophil-to-leucocyte ratio (dNLR), and (c) platelet-to-lymphocyte ratio (PLR). \textit{URun} refers to unrestricted cross-validation, \textit{RRun} refers to restricted cross validation.
4. Discussion

4.1. Summary of Main Findings

Cross-validation with random split of the data set under restrictions, namely respecting the proportion of those with and without metastasis, showed that the ROC classification models of the investigated ratios had low variability reported to the full mode. The variability of ROC classification models in the unrestricted samples is very similar to the full model, when almost 70% of cases are in the development set with identical cut-off for 4/5 ratios (excepting PLR). For all other unrestricted samples, higher variability of the AUC was observed as compared to the restricted samples (Figure 2, Figures S1–S3).

4.2. Findings Discussion

The cut-off values are different, when different methods are used (Table 1), as expected and this affects the number of absolute frequencies in the confusion matrix. The Youden index maximizes both sensitivity and specificity and retrievs higher cut-off values for our dataset, as compared to minimization of the distance to the point with (0,1) coordinates. The number of false positives increases and the number of false negatives decreases, when the cut-off values are identified by the Tanget method as compare with the Youden’s index, which reflectes on both Se (increases) and Sp (decreases), but without significant effects on the LRs or CUIs (Table 2). Even if the Tangent method showed superiority in identification of the true cut-off value [58], this was not observed in our sample. Despite the method used to define the threshold values, there are very poor case-findings. NLR and dNLR are fair for screening, while PLR is poor for screening (Table 2). On the other hand, they are not recommended as biomarkers for metastasis in patients with colorectal cancer.

As expected, different cut-off values were observed with few exceptions in both restricted and unrestricted cross-validation approaches (Tables 1, 3 and 4). Generally, not only for the unrestricted, but also for the restricted cross-validation, the cut-off values were lower than those of the full model. It resulted exactly the opposite for dNLR and PLR. Different cut-off values reported in cross-validation (Tables 3 and 4) indicate a variability, most probably caused by the characteristics of the included observations (the input data) and it has been observed in previously reported researches on patients with colorectal cancer [47–50].

The characteristics of the models were similar for both scenarios, with the AUCs values confined in the confidence interval of the AUC of the full model, without exception and with the AUC of the full model confined in the confidence interval of each individual classification model, regardless of the approach applied (Tables 1, 3 and 4). Opposite significances of the AUCs in unrestricted cross-validation in comparison to the full model was observed for PLR (2/13 classification models, one model that included 35.4% subject from the cohort and the other one that included 15.4%). The consistency of the significance across different samples indicates the stability of the classification models for the investigated ratios and the presence of metastasis as the outcome.

The cut-off values for 4/5 ratios were identical with those of the full model for the first unrestricted run, with almost 70% of subjects randomly selected for the development set. The relative value of the cut-off value from the full model ((Cut-Off_{model}−Cut-Off_{full-model})/Cut-Off_{full-model} × 100) was generally lower for the restricted models as compared to the unrestricted model. It can be explained by the small differences for those runs with ~70% of observations in the development sets. A pattern is observed, when the cut-off values are investigated. There are lower cut-off values in the majority of the cases for NLR (61.5% in the unrestricted and 80% in the restricted) and higher cut-off values for dNLR (61.5% in the unrestricted runs and 100% in the restricted runs) or PLR (53.8% in the unrestricted runs and 80% in the restricted runs). The cut-off values are reflected in the number of TP, TN, FN and FN and thus in the performances of the classification model (Tables 2 and 6). The evaluation of the full model performances showed that none of the investigated ratios is a good predictor for metastasis on patients with colorectal cancer (Table 2), with low +LR, high -LR, a relatively high
misclassification rate and MCC values less than 0.25. The low performances of the classification model are also observed on restricted cross-validation (Table 6), supporting the results reported by Martens et al. [59], namely that small changes in AUCs bring small changes in performances of the classification model. The ROC analysis is used to overcome this shortcoming, namely poor correlation of one predictor with a certain outcome, in order to find the threshold for classification and, based on the threshold, to consecutively identify a multivariable predictive model [42–50,60,61]. A small number of articles reported the threshold for the investigated inflammatory ratios on patients with colorectal cancer regarding metastasis as the outcome. Anuk and Yıldırım reported a cut-off value for PLR equal to 194.7 for liver metastasis (Se = 74.5%, Sp = 72.7%) and of 163.95 (Se = 56.8%, Sp = 56.3%) for lymph node tumour cells invasion [62] on a sample of 152 patients. The thresholds for PLR reported in this study are higher than those reported by Anuk and Yıldırım [62], nonetheless the specificity is higher at the expense of sensitivity (Tables 2 and 6). In our study we did not devided into the type of metastasis, but the differences in thresholds and Se(s) are important and could be possible explained by the characteristics of the target population. Guo et al. showed a large variability of PLR (25.4 to 300) thresholds in the evaluation of survival of the patients with colorectal cancer [45]. High PLR was associated with poor overall survival and high LMR was associated with favourable overall survival. The different cut-off values reported by Guo et al. should be explain by different population included in the analysis of the primary researches [45]. This explanation is also supported by the results reported by Xia et al. on a sample of 154 patients with T1-2 rectal cancer with cut-off of 2.8 for NLR and of 140.0 for PLR, when the primary outcome was considered death and the secondary one it was the occurrence of postoperative complications [46]. Peng et al. reported the optimal cut-off values of 4.63 for NLR (AUC = 0.568, 95%CI (0.477 to 0.660), p = 0.152) and 150.17 for PLR (AUC = 0.518, 95%CI (0.424 to 0.611), p = 0.707) on patients with colorectal cancer liver metastasis only [63]. The optimal cut-off values reported in this study are different by the those previously reported in the scientific literature. However, larger studies are needed to find the explanation for the reported thresholds and models performances.

The use of area under the ROC curve (AUC) in the assessment of classification model performances is known to be biased especially for small samples [64–68] and different cross-validation approaches have been proposed for model validation. The cross-validation approaches are usually used when computer algorithms are applied to identify the best performing classification model [64–69]. Moreover, leave-pair-out (LPO) cross-validation showed low bias in AUC estimation [70]. In our study, we applied cross-validation outside the computer algorithm in identification of the classification model and the unrestricted random sample with 70% of observations in the development set and it proved similar performances with the restricted cross-validation also with 70% of observations in the development set (Tables 1–3 and 5, Figures S1–S3). Our result supports thus the conclusion that, for large samples, an appropriate random sample with inclusion in the study of 70% of eligible subjects will closely reflect the target available population. Similarly, for large samples, the experimental design is the key factor for a valid ROC classification model, so the dictum “Garbage in, garbage out” seems true even for large data sets [71].

4.3. Study Limitations

Despite of a rigorous experimental design we need to list some limitations of our study. First limitation refers to the cross-validation methods used. The unrestricted cross-validation method was applied using the options of the program and unfortunately, the split was not saved, thus the performances of the models could not be reported. The simple random sampling proved its ability of appropriate splitting of a dataset in development (training) and validation (test) set, in the context of a predictive linear multivariate regression model [72]. Furthermore, only two cross-validation methods were applied and the use of other cross-validation methods (e.g., minimum of $p$-value on the confusion matrix [73], max($Se \times Sp$) [74], min($Se − Sp$) [75]) could retrieve different cut-off values and thus different performances of classification models. However, since simple random
splitting in development and validation sets with ~70% of the observations in the development set and the restricted cross-validation (that also included 70% of the observations in the development sets) perform closely to the full model, no significant changes in terms of clinical utility are expected. Second limitation refers to the number of runs for each scenario, a higher number of runs being able to better reflect the reality. As it could be observed, we decided to report the performances in individual runs, not only to present an average of the performances, which is a common practice, but also to closely monitor at what length the performances differ from each other. Nevertheless, since the classification models showed stability in both restricted and unrestricted cross-validations, we are not expecting significant add-on with the increase in the number of runs. Third limitation is related to the considered outcome. We did not separately evaluated different types of colorectal cancer metastasis (e.g., liver, lung, brain, peritoneum, distant lymph nodes) and it can be expected that the thresholds of the investigated ratios to vary with the type of metastasis. However, for such evaluation we need to expand the time frame for the retrospective evaluation, in order to assure a sufficient number of observations for each type of metastasis. Fourth limitation is related to the investigated inflammatory ratios. Investigation of the behaviour of other inflammatory ratios, such as lymphocyte-to-monocyte ratio, systemic immune inflammation index or prognostic nutritional index is also of interest. All these markers are taken into consideration by our research team and they will be tested for prediction abilities and proof of clinical utility.

5. Conclusions

The investigated ratios proved low clinical utility in predicting metastasis on patients with colorectal cancer. The full models showed fair clinical utility in screening, but the values of positive and negative likelihood ratios did not support their application in clinical settings. The classification models identified in the development sets, regardless of the use of unrestricted or restricted (percentage of patients with and whithout metastasis) random split, showed characteristics and performances similar to the full models. Our results showed that a simple random split of observations, weighting or not the patients with and whithout metastasis, in a ROC analysis assures the performances similar to the full models when at least 70% of the available population is included in the evaluation. Moreover, the sample size of the development set in the case of ROC classification analysis of investigated inflammatory ratios, considered as predictors for metastasis on patients with colorectal cancer, had little interest, in most of the cases, either on the models characteristics or on the models performances. The cut-off values of the investigated scenarios should probably be explained by the characteristics of the input data and are, most likely, linked to the percentage of correct and false classification, but without a significant impact on the models characteristics or performances. This behaviour supports the stability of the classification models in the context of a proper experimental design.

Supplementary Materials: The following are available online at [http://www.mdpi.com/2227-7390/8/10/1741/s1](http://www.mdpi.com/2227-7390/8/10/1741/s1), Figure S1: “AUCs of neutrophil-to-lymphocyte ratio (NLR) in unrestricted (URun) and restricted (RRun) cross-validation related to the full model”, Figure S2: “AUCs of derived neutrophil-to-leucocyte ratio (dNLR) on unrestricted (URun) and restricted (RRun) cross-validation related to the full model.”, Figure S3: “AUCs of platelet-to-lymphocyte ratio (PLR) on unrestricted (URun) and restricted (RRun) cross-validation related to the full model”.  

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