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To cite this version:

Dimitri Rochman, Eric Baugé, Alexander Vasiliev, Hakim Ferroukhi, Gregory Perret. Nuclear data correlation between different isotopes via integral information. EPJ N - Nuclear Sciences & Technologies, EDP Sciences, 2018, 4, pp.7. 10.1051/epjn/2018006. cea-01794006

HAL Id: cea-01794006
https://hal-cea.archives-ouvertes.fr/cea-01794006
Submitted on 2 Oct 2019

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Nuclear data correlation between different isotopes via integral information

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Received: 15 September 2017 / Received in final form: 24 January 2018 / Accepted: 19 March 2018

Abstract. This paper presents a Bayesian approach based on integral experiments to create correlations between different isotopes which do not appear with differential data. A simple Bayesian set of equations is presented with random nuclear data, similarly to the usual methods applied with differential data. As a consequence, updated nuclear data (cross sections, fission neutron spectra and covariance matrices) are obtained, leading to better integral results. An example for 235U and 238U is proposed taking into account the Bigten criticality benchmark.

1 Introduction

It was recently demonstrated that an uncertainty decrease and non-zero correlation terms between different nuclear data reactions can be obtained when using integral information such as criticality benchmarks [1] (see Refs. [2–4] for other examples). In reference [1], cross-correlation terms between \( \Phi \) (emitted neutrons per fission), \( \chi \) (fission neutron spectra) and \( \sigma_{(n,f)} \) (fission cross section) were calculated in the case of the \(^{239}\text{Pu}\) isotope with specific Pu benchmarks in the fast neutron range. Such approach can be useful to lower calculated uncertainties on integral quantities based on nuclear data covariance matrices, without artificially decreasing cross section uncertainties below reasonable and unjustified values. This is appropriate when the propagation of uncertainties from differential data to large-scale systems indicates an apparent discrepancies between uncertainties on measured integral data (neutron multiplication factor, boron concentration, isotopic contents) and the calculated ones. In this reference, the correlation terms between reactions for a specific isotope and the decrease of differential uncertainties were calculated using a simple Bayesian Monte Carlo method. In the present work, the same method is applied (1) to obtain correlation terms this time between different isotopes, and (2) to decrease the uncertainties for important reactions, using again criticality-safety benchmarks. The approach and the equations used in the present work are the same as in [1].

In the following, the case of the \(^{235}\text{U}\) and \(^{238}\text{U}\) isotopes will be considered and the Bayesian update will be performed using a specific criticality benchmark with high sensitivity to these isotopes: the intermediate metal fast number 7 benchmark, or \text{imf}7\footnote{Also known as Bigten [5].} First the method will be recalled in simple terms, then the application with the \text{imf}7\footnote{Benchmark} will be presented. The updated benchmark value, cross sections, correlations and uncertainties will be compared to the prior values, thus demonstrating the results for the differential quantities. This is of interest in the context of nuclear data evaluations, where both nominal values and covariance matrices can reflect the present results.

2 Correlation from integral benchmarks

The basic principles of the method were already presented in [1]. We will outline here the major equations. The Bayesian updates of the prior information is obtained using a Monte Carlo process:
- random nuclear data are produced following specific probability density functions (pdf). Such pdf were obtained as follows: starting from uniform distributions, comparisons between calculations and differential measurements (from EXFOR) were performed. Following the description of reference [6] (and as presented below for integral data), weights are derived from such comparisons and pdf of TALYS model parameters are updated. The next step is to sample from these specific parameter pdf to produce random nuclear data;
- each random nuclear data is used in the benchmark simulation;
the random calculated quantities are compared to the measured one, and;
- finally each random nuclear data is weighted according to the agreement between the calculated and measured quantities (see below for details on the definition of such weights).

In the present work, the \( k_{\text{eff}} \) value of the imf7 benchmark is used as the only integral quantity: the reported value in [5] is \( k_{\text{exp}} = 1.00450 \) with an experimental uncertainty of \( \Delta k = 70 \text{ pcm} \). As a prior for the nuclear data, the random \( ^{235}\text{U} \) and \( ^{238}\text{U} \) cross sections (and emitted particles and spectra) are obtained from the TENDL-2014 library [7]. The T6 system [8] was used to generate so-called random ENDF-6 and ACE files, containing all necessary random nuclear data. This way, the same file production and processing is followed, based on TALYS and NJOY [8,9]. In the case of the imf7 benchmark, the \( k_{\text{eff}} \) value is very sensitive to the unresolved resonance range [10] and the ENDF-6 files are processed with the PURLR module of NJOY. Each ENDF-6 and ACE files are similar in format, but different in content. They are based on sampling of model parameters of the different nuclear models according to specific independent probability distributions (see the TMC, BMC, UMC-B and BFMC methods [6], [11–13] for details). Model parameters are sampled a large number of times (with the index \( i = 1 \ldots n \)) to generate full cross sections and other nuclear data quantities for \( ^{235}\text{U} \) and \( ^{238}\text{U} \) from 0 to 20 MeV (see for instance [14] for the testing of such file distributions). The sampling between these two isotopes is performed in independent manner, so that no correlation between \( ^{235}\text{U} \) and \( ^{238}\text{U} \) can exist other than from the model themselves. The prior correlation matrices for \( ^{235}\text{U} \) and \( ^{238}\text{U} \) are simply obtained from the \( n \) random files, using the conventional covariance and standard deviation formula.

The \( n \) random ACE files are then used in \( n \) MCNP6 simulations [15], leading to \( n \) values of calculated \( k_{\text{eff},i} \), with \( i \) varying from 1 to \( n \). The comparison between \( n \) random calculated \( k_{\text{eff},i}=1 \ldots n \) and the experimental value \( k_{\text{exp}} \) is performed with the simplified chi-2 \( Q \) values and associated weights \( w_i \) (here, chi-2 is called \( Q \), to differentiate it from the neutron spectra \( \chi \)):

\[
Q_i = \left( \frac{k_{\text{eff},i} - k_{\text{exp}}}{\Delta k} \right)^2 \tag{1}
\]

\[
w_i = \exp \left( -\frac{Q_i}{2} \right). \tag{2}
\]

Such formulation can easily be linked to the usual Bayesian likelihood [13,16]. The weights are then assigned to the corresponding \( ^{235}\text{U} \) and \( ^{238}\text{U} \) nuclear data files (for both isotopes together) which lead to \( k_{\text{eff},i} \). Considering \( n \) random files for each isotope, there is \( n^2 \) possible combinations; in the following, we will consider only \( n \) combinations such as (1,1), (2,2),...,(\( i,i \)).

Examples for the weights of the random \( ^{235}\text{U} \) and \( ^{238}\text{U} \) nuclear data are presented in Figure 1. In this example, one iteration \( i \) corresponds to the use of one specific random file for \( ^{235}\text{U} \) and another one for \( ^{238}\text{U} \). As observed, the distribution of the weights \( w_i \) strongly varies from values close to 1 (for \( Q_i \approx 0 \), indicating a good performance of the random files \( i \)) to very small values (almost 0 for large discrepancies between \( k_{\text{exp}} \) and \( k_{\text{eff},i} \)). Due to this large range of weights, a large number of random files is necessary to obtain meaningful results. In the case of 7000 random files for each U isotope, about 18% of the weights are higher than 0.01.

The final quantity for a specific benchmark consists of a matrix containing \([i, \sigma_i(^{235}\text{U}), \sigma_i(^{238}\text{U}), w_i]\) for \( i = 1 \ldots n \), where \( \sigma_i \) stands for all nuclear data quantities as a function of energy. As previously mentioned, the value of \( n = 7000 \) is considered in this work. The correlation \( \rho(\sigma_a, \sigma_b) \) can be calculated for specific values of the incident neutron energies for \( \sigma_a (E_k) \) and \( \sigma_b (E_k) \). For instance, \( \sigma_a \) is the fission cross section of \( ^{235}\text{U} \) and \( \sigma_b \) is the capture cross section of \( ^{238}\text{U} \), both at a specific energy \( E_k \) and \( E_{k'} \), respectively. Considering the vector \([i, \sigma_i(^{235}\text{U}), \sigma_i(^{238}\text{U}), w_i]\), \( \rho \) can be calculated as follows. Using the definition of weighted averages:

\[
\begin{align*}
\omega &= \sum_{i=1}^{n} w_i \\
\omega_{\alpha\beta} &= \sum_{i=1}^{n} w_i \sigma_{\alpha,i} \sigma_{\beta,i} / \omega
\end{align*}
\]

and the definition of the weighted variance/covariance factors:

\[
\begin{align*}
\text{var}_{\sigma_{\alpha}} &= \sum_{i=1}^{n} \frac{[\sigma_{\alpha,i} - \omega_{\alpha} \cdot w_i]^2}{\omega} \\
\text{var}_{\sigma_{\beta}} &= \sum_{i=1}^{n} \frac{[\sigma_{\beta,i} - \omega_{\beta} \cdot w_i]^2}{\omega} \\
\text{cov}_{\sigma_{\alpha}\sigma_{\beta}} &= \sum_{i=1}^{n} \frac{[\sigma_{\alpha,i} - \omega_{\alpha} \cdot w_i \sigma_{\beta,i} - \omega_{\beta} \cdot w_i \sigma_{\alpha,i}]}{\omega}
\end{align*}
\]
The correlation \( r(s_a, s_b) \) between \( s_a \) and \( s_b \) is given by
\[
 r(s_a, s_b) = \frac{\text{cov}(s_a, s_b)}{\sqrt{\text{var}(s_a) \cdot \text{var}(s_b)}}.
\] (3)

Such correlation \( r \) can be obtained for different \( E_R \) and \( E_p \), thus defining a full correlation matrix between the same cross section and the same isotope, between different cross sections for the same isotopes, and between isotopes. As quantities in these equations (average cross sections, standard deviations and correlation factors) come from a Monte Carlo process, one has to check their convergence as a function of the iteration number, as presented in Figure 2.

One can see that in both cases (considering or not weights \( w_i \)), the final correlation values are different, and the difference is outside the standard errors (defined as \( \sqrt{\frac{r_i^2}{n-2}} \) for the non weighted case). As it can be seen on this figure, the non weighted running correlation evolves smoothly with the increasing number of samples, while the weighted running correlation exhibits large jumps for low iteration \( i \) where high weight samples are added to the calculation (as seen in [16] showing same kind of behavior).

In the following, more details will be given on the imf7 benchmark together with the results regarding the prior and posterior information for the uranium isotopes.

### 3 Application to \(^{235}\text{U} \) and \(^{238}\text{U} \)

The work presented in [1] was limited to the single \(^{239}\text{Pu} \) isotope, since it was applied to integral experiments from the PMF subtype (Plutonium Metal Fast) of the ICSBEP collection [5], for which only \(^{239}\text{Pu} \) nuclear data dominate the benchmark calculation result. Following the same idea, the imf7 benchmark is selected as its \( k_{	ext{eff}} \) is highly impacted by both \(^{235}\text{U} \) and \(^{238}\text{U} \).

#### 3.1 The imf7 Benchmark

The imf7 benchmark (intermediate enrichment uranium metallic fast number 7), also known as Bigten, is a highly enriched uranium core, surrounded by a massive natural uranium reflector. It is characterized as a fast system, as the majority of the neutron spectrum is above 100 keV. Bigten is a cylindrical assembly with a core composed entirely of fissionable material in metal form. There are three distinct regions: a nearly homogeneous cylindrical central core made of uranium enriched at 10% in \(^{235}\text{U} \),

![Fig. 3. Neutron spectrum of the imf7 ICSBEP benchmark calculated by MCNP6 using TENDL-2014 nuclear data. This spectrum is averaged for the whole benchmark. In the \(^{238}\text{U} \) blanket, the average neutron energy is 345 keV, while in the \(^{235}\text{U} \) core, it is 580 keV.](image-url)
surrounded by a heterogeneous core volume made of natural uranium and highly enriched uranium (93%) and a cylindrical reflector, made of depleted uranium, completely surrounding the core. Figure 3 shows the neutron spectrum averaged over imf7, calculated using MCNP6 with TENDL-14 nuclear data, and average energies for fission and capture are presented in Table 1. It has a typical fast spectrum with an average neutron energy of 530 keV.

This imf7 configuration has long been known by evaluators to be sensitive to nuclear data for both $^{235}\text{U}$ and $^{238}\text{U}$ isotopes. This double dependency is so strong that mixing nuclear data for $^{235}\text{U}$ from one source (e.g. ENDF/B-VII.1 [17]) with data for $^{238}\text{U}$ from another source (e.g. JEFF-3.3) in a imf7 benchmark calculation, results in a poor restitution of the measured $k_{\text{eff}}$ value. Some examples are presented in Table 2 by repeating the benchmark calculation with different nuclear data evaluations for $^{235}\text{U}$ and $^{238}\text{U}$.

As observed, if both uranium isotopes come from the same library, the calculated $k_{\text{eff}}$ is close to the experimental value. On the other hand, a mixture of the library of origin leads to very different calculated $k_{\text{eff}}$. These cases can be interpreted as the effective presence of correlated isotopes in current evaluated nuclear data libraries.

### 3.2 Correlations

By extending the methodology described in reference [1], such cross-isotopes correlations can be rigorously quantified. All combinations of neutron incident energy, observables (cross sections, prompt fission neutron spectra, nubar, etc.), and target isotopes are possible, as illustrated in Figure 4.

Correlation matrices for a selection of cross sections, nubar and pfns in the case of $^{235}\text{U}$ and $^{238}\text{U}$. Top: correlation without taking into account the imf7 benchmark; bottom: same, but taking into account imf7. See text for details. In each sub-block, the cross sections are presented as a function of the incident neutron energy (the lower-left part corresponds to the lower neutron energy range, whereas the higher-right part corresponds to the higher neutron energy).

The upper panel of Figure 4 shows the full $^{235}\text{U}$-$^{238}\text{U}$ correlation matrix for the prior (unweighted), Total Monte-Carlo (TMC) [11] samples for $^{235}\text{U}$ and $^{238}\text{U}$, as computed from the TENDL-2014 library. Four blocks are separated by two red lines, each block represents the correlation and cross-correlation for these isotopes: bottom-left: $^{235}\text{U}$-$^{235}\text{U}$, bottom-right: $^{235}\text{U}$-$^{238}\text{U}$, top-left: $^{238}\text{U}$-$^{235}\text{U}$ and top-right: $^{238}\text{U}$-$^{238}\text{U}$. As it can be seen, cross-isotopes correlations between isotopes are zero, since model parameters for both isotopes were independently sampled in this study.

![Fig. 4.](image)

The lower panel shows the full $^{235}\text{U}$-$^{238}\text{U}$ correlation matrix for the TMC samples of $^{235}\text{U}$ and $^{238}\text{U}$, weighted according to equation (2), where $k_{\text{exp}}$ is the experimental value of the imf7 benchmark, and $k_{\text{eff,i}}$ that derived from the $^{235}\text{U}$ and $^{238}\text{U}$ sampled files, indexed by $i$. Obviously, that lower panel exhibits cross-isotopes correlations contrary to the upper one, and it also exhibits correlations between different types of observables like those discussed in [1].

Although the TMC treatment allows the constructions of covariance matrices between all the nuclear data observables, the matrices shown in Figure 4 are restricted to the observables which are expected to have a strong influence of $k_{\text{eff}}$, hence the $(n,p)$, $(n,2n)$, and other cross sections are not shown in this figure. The color coding of the amplitude of the correlation in Figure 4 reflects four levels of correlations: zero or very low (white), low (lighter blue or red), moderately strong (intermediate blue or red), and very strong (darker blue or red), with red identifying positive correlations, and blue negative ones. The correlations between observables from different isotopes (in the off-diagonal blocks) sit in the low range. The $^{235}\text{U}$ or $^{238}\text{U}$...
sub-matrices display some stronger correlations, mostly along the diagonal, but also for observables derived from the optical model potential (total, non elastic and elastic cross sections), highlighting the role played by that model in inducing correlations in nuclear data.

As expected, similarly to the conclusions of references [1,16], a weak negative correlation for the posterior is observed (see Fig. 5 for an enlarged sub-matrix) between the \( \sigma_{(n,f)} \) of \( ^{235}\text{U} \) and its fission cross section, for energies close to the mean energy of neutrons causing fission in \( ^{235}\text{U} \) (Tab. 1). This anti-correlation results from \( \sigma_{(n,f)} \) being two factors in the product describing the neutron source term in the neutronic transport equation: a stronger \( \sigma_{(n,f)} \) is exactly compensated by a weaker \( \sigma \).

The correlation matrix between the \( ^{235}\text{U} \) capture and fission cross sections (Fig. 6) is harder to interpret, since it exhibits a complex structure. Although the crosses materializing the mean energies leading to fission and capture reactions in the core and blanket regions of the assembly both sit in the weak correlation region of the map (close to the negligible correlations zone (white), there are regions of stronger correlation, both positive and negative, nearby. The moderate positive correlation for neutron energies seen above 500 keV can be understood as \( ^{235}\text{U}(n,f) \) driving the source term of the neutronic transport equation and \( ^{235}\text{U} \) \((n,\gamma)\) being a contributor to the absorption term of that equation. For lower neutron energies, two zones of moderate negative correlation are observed, one for low \( (E < 200\text{keV}) \) neutron energy inducing fission, and one for low neutron energy inducing capture. That complex structure of the \( ^{235}\text{U} \) capture and fission correlation might result from the interplay between \( ^{235}\text{U} \) in the core region (fast spectrum) and the blanket region (slower neutronic spectrum).

From Figure 4, one can also note two important aspects:
- anti-correlation for \( ^{235}\text{U} \) between \( \gamma \) and \((n,\gamma)\): in order to compensate for a higher neutron capture, the fission spectrum becomes harder, thus producing more neutrons at higher energy;
- especially in the case of \( ^{238}\text{U} \), anti-correlation appears in the updated matrices between the inelastic cross sections themselves. Again, this can be understood in order to compensate for the loss of neutrons caused from a specific inelastic cross section (for instance \((n,\text{inl})\)) by another one (for instance \((n,\text{inl2})\)).

In the off-diagonal cross-isotope correlation blocks, a prevalent weak positive correlations can be observed between \( ^{235}\text{U}(n,f) \) and \( ^{238}\text{U}(n,\gamma) \) at energies where the neutronic spectrum is strong (see Fig. 7 for an enlarged sub-matrix). Again, that positive correlation is explained by \( ^{235}\text{U}(n,f) \) driving the source term and \( ^{238}\text{U}(n,\gamma) \) being the other strong contributor to the absorption term of the neutronic transport equation.

A very prevalent weak anti-correlation can also be observed between the fission cross section of \( ^{235}\text{U} \) and the total elastic cross section of \( ^{238}\text{U} \) (presented in an enlarged format in Fig. 8). They are anti-correlated since a weaker fission cross section of \( ^{235}\text{U} \) can be compensated by a more efficient neutron reflector \( ^{238}\text{U}(n,\text{el}) \), which reflects leaking neutrons back into the \( ^{235}\text{U} \) core for another attempt to fission \( ^{235}\text{U} \).

### 3.3 Updated cross sections and variances

The weighting of TMC samples according to equations (1) and (2) not only introduces correlations between observables, but it also leads to modifications of the central values
of nuclear data as well as a reduction of the variances of the various nuclear data observables. Such updated cross sections and variances are presented in Figures 9 and 10 and for all considered quantities.

Ratio of cross sections (and $\nu$ and $\chi$) for the post-adjusted (a posteriori) over the prior. The cross sections, $\nu$ and $\chi$ are presented from 100 keV to 6 MeV on a logarithmic scale.

The general observation is that the cross sections (including $\nu$ and $\chi$) are moderately updated (maximum of 1.0% for the $235\text{U}(n,\text{f})$ cross section) whereas the variances are strongly reduced (see for instance $237\text{U}(n,\text{f})$). The changes in the posterior cross sections are to some extent depending on the prior uncertainties. If the prior uncertainties are small, the changes will also be small. Therefore the changes presented in Figure 10 can be different for different prior. In the case of $238\text{U}$, that reduction brings the variance in the same order of magnitude as that of the existing experimental differential data. However, for $238\text{U}$, the reduced standard deviation is still larger than that of existing differential data: a further Bayesian update with that differential data would further reduce the calculated uncertainty of the $238\text{U}$ $\nu$ (see for instance [18,19] for details).

A limited set of cross section uncertainties is strongly affected by the Bayesian update: with a decrease for $235\text{U}(n,f)$, $235\text{U}(n,\text{inl}), 234\text{U}(n,\gamma)$, $238\text{U}(n,\text{inl})$, $235\text{U}(n,\text{el})$ and an increase for $238\text{U}(n,\text{inl})$. One should notice that the $(n,\text{inl})$ cross section for $238\text{U}$ is relatively small, with a maximum at 400 mb, compared to the $(n,\text{inl})$ cross section (with a maximum of 1.5 b). Such change could be explained by statistical fluctuations, but a dedicated study on this effect would be necessary to clarify its origin. The increase of this cross section uncertainty has therefore a limited impact. It is difficult to assess the relative importance of these cross sections in the decrease of the $k_{\text{eff}}$ uncertainty, but the mentioned reactions are important for the account of neutrons in the energy region of interest.

The value of the $235\text{U}$ posterior fission cross section is modified by a factor as large as 1.003 relatively to that of the prior, and its standard deviation is strongly reduced. When compared with the international cross section standard [19] for the $235\text{U}$ fission cross section (see Fig. 11), their agreement is quite good over an extended energy range: the central values are close (except after the onset of the second chance fission around 0.8 MeV, where the posterior cross section overestimates that of the standard) and the error bars largely overlap. For $238\text{U}$, the relative variations of the posterior with respect to the prior are less than 1%.

As a final remark, since the Bayesian weighting of samples applies to sets of complete ENDF-6 formatted files (one set including an ENDF-6 file for $235\text{U}$ and a file for $238\text{U}$), that weighting process produces adjustments and variance reduction for all the observables included in these files, from the inelastic and elastic cross sections, which do play role in the calculation of imf7, to cross sections like $(n,p)$ or $(n,\alpha)$, which are hardly constrained by the benchmark.

### 3.4 Resulting $k_{\text{eff}}$ distributions

Table 2 shows the averages and standard deviations of the calculated $k_{\text{eff}}$ distributions, compared with the experimental value, with unweighted samples labeled as “prior”, and weighted samples labeled as “posterior”. Those distributions of $k_{\text{eff}}$ are also displayed on Figure 12. In Table 3 and Figure 12, the posterior distribution can be observed to agree very well with the experimental result and its uncertainties, while the average $k_{\text{eff}}$ resulting from the unweighted prior is lower, with a much wider distribution.

### 4 Discussions

As mentioned in the introduction, the goal of this type of work is to reduce the calculated uncertainties on integral quantities while keeping realistic uncertainties and correlations for the differential data. Additionally, as shown in Table 3 for imf7, the updated $235\text{U}$ and $235\text{U}$ nuclear data provide $k_{\text{eff}}$ which is in better agreement with the experimental value. Such method can be extended by including more benchmarks in the definition of $Q_i$ (and also by including other quantities such as spectra indexes), but prior to the continuation, two tests can be performed. The first one is partially presented in Figures 9 and 10, showing that the
updated nuclear data are still in agreement with the differential data (i.e. pointwise cross sections, or pointwise $\sigma$). This is not explicitly shown in these figures, but the fact that the updated cross sections are very close to the prior values indicates that the method does not produce very different cross sections compared to the prior. And as it was mentioned, the agreement with the standard cross section is still respected, given the large variances of the TENDL curves.

The second test concerns the predictive power of the method: by choosing a benchmark with similar characteristics than imf7, is its calculated $k_{\text{eff}}$ improved? If this is the case, one can consider that the indications provided by the updated cross sections are general enough to be exported to outside the case of imf7. To answer this question, three additional benchmarks are calculated with the same random $^{238}\text{U}$ and $^{235}\text{U}$ nuclear data files; using or not the weights from imf7. Two of these benchmarks are relatively close to imf7:

Fig. 9. Ratio of cross sections (and $\sigma$ and $\chi$) for the post-adjusted (a posteriori) over the prior. The cross sections, $\sigma$ and $\chi$ are presented from 100 keV to 6 MeV on a logarithmic scale.

Fig. 10. Same as Figure 9 but for the calculated uncertainties (standard deviations).
A third benchmark is on purpose chosen to be very different than imf7: it is a thermal system of low-enriches UO fuel rods with a high water-to-fuel ratio: lct6-1. For this benchmark, the modifications of the $^{238}$U and $^{235}$U nuclear data in the fast neutron range from imf7 are expected to have little impacts on the calculated $k_{\text{eff}}$. The results of these calculations are presented in Table 3.

First, the $k_{\text{eff}}$ values for hmf1, imf1-1 and lct6-1 calculated with weights from imf7 (posterior in Tab. 3), are not in worse agreement with experiment than the ones calculated without weights (prior in Tab. 2). This suggests that weighting random samples according to one given benchmark does not produce a distribution that is only good for that benchmark. Moreover, introducing the imf7-derived weights seems to slightly improve the agreement of all three of our test cases with experimental values, suggesting that the changes due to that weighting carry some real physics and are not just a better local optimization. However, while the weighted imf1-1 and lct6-1 calculation results are within experimental uncertainties, that of hmf1 is still well outside of experimental uncertainties, suggesting that the imf7 specific weighting is missing some of the physics that is essential for the hmf1 case.

Now, looking at the calculated uncertainties for the weighted hmf1 and imf1-1 cases, we observe that their widths are reduced compared to those of the unweighted calculations, suggesting again that imf7-derived weights carry some real physical information. However, the widths resulting from weighted calculations are much larger than experimental uncertainties. In the case of the lct6-1 benchmark, the uncertainties are not reduced: the changes generated at high energy do not impact the uncertainties for this thermal system. This indicates that in the case of a general evaluation of nuclear data, one needs to include benchmarks spanning over a wide energy range.

In order to confirm the conclusions from the above test, it should be repeated on a more extensive set of benchmark cases. The next step in this process would then be to calculate weights from all those benchmark cases, to combine them (maybe through a simple product), and test whether the resulting weighted distribution provides a good restitution of all the experimental benchmark data used to determine those weights (see for instance the work performed in [3, 20]).

There is also no reason to restrict the benchmark data used to calculate weights to only $k_{\text{eff}}$, and other types of data, like spectral indices or differential measurements, are likely to carry information that constrains nuclear data in a different manner.

5 Conclusion

It has been shown that including integral constraints from experiments that are sensitive to two isotopes introduces effective cross-correlations between the nuclear data of
these isotopes. It was demonstrated that it is possible to quantify such cross-correlation between isotopes using an integral benchmark, based on a Bayesian method and a set of random nuclear data. The case under study concerns the $^{235}$U and $^{238}$U isotopes and the Bigten (im7) benchmark. Additionally, the updated nuclear data and their covariance matrices lead to a better agreement with the calculated and measured integral data, for the central values and for uncertainties, while keeping the original good agreement with differential data.

This is an extension of the method previously proposed for $^{239}$Pu [1] and is a confirmation that such method allows (1) to be part of the evaluation process of nuclear data, and (2) to obtain reasonable integral and differential uncertainties. In the future, the method will be applied taking into account a larger set of integral data and exploring applications below the fast neutron range. Our limited testing is suggesting that weighting with respect to one benchmark experiment does not negatively affect the agreements with other experiments and even improves them slightly. A more extensive testing is needed to confirm that combining weights calculated from different benchmark experiments leads to a weighted sampling that simultaneously accounts for all those benchmarks and their associated uncertainties. Such a combination of weights originating from different benchmarks will be the subject of a forthcoming article.

Like in [1], the present work is at the “proof of concept” stage: the methodology seems to work with a reduced set of integral constraints and the rather simple models used to produce the TENDL-2014 library. In order to produce evaluations of the quality of the best evaluated nuclear data libraries, that method will have to be extended to: 
- include a larger and more representative set of integral experimental constraints, spanning a wide range of neutronic spectra and applications;
- include integral constraints other than $k_{\text{eff}}$ in the calculation of weights, include differential constraints as well as international cross sections standards [19] in the calculations of weights;
- apply that methodology to the more sophisticated models [21–23] used to evaluate the nuclear data of the best international data libraries;
- completely implementing the above extensions would produce fully updated nuclear data and covariance matrices, including cross-isotopes and cross-observables correlations, following a well defined reproducible scheme. These files should allow for accurate simulation of application, including calculated uncertainties. Such work would then be part of the elaboration of a nuclear data library based on models (for differential data), realistic model parameter distributions and integral constraints, as presented in [24].

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Cite this article as: Dimitri A. Rochman, Eric Bauge, Alexander Vasiliev, Hakim Ferroukhi, Gregory Perret, Nuclear data correlation between different isotopes via integral information, EPJ Nuclear Sci. Technol. **4**, 7 (2018)