External validation of a magnetic resonance imaging-based algorithm for prediction of side-specific extracapsular extension in prostate cancer

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Introduction
Recently developed algorithm for prediction of side-specific extracapsular extension (ECE) of prostate cancer required validation before being recommended to use. The algorithm assumed that ECE on a particular side was not likely with same side maximum tumor diameter (MTD) <15 mm AND cancerous tissue in ipsilateral biopsy <15% AND PSA <20 ng/mL (both sides condition). The aim of the study was to validate this predictive tool in patients from another department.

Material and methods
Data of 154 consecutive patients (308 prostatic lateral lobes) were used for validation. Predictive factors chosen in the development set of patients were assessed together with other preoperative parameters using logistic regression to check for their significance. Sensitivity, specificity, negative and positive predictive values were calculated for bootstrapped risk-stratified validation dataset.

Results
Validation cohort did not differ significantly from development cohort regarding PSA, PSA density, Gleason score (GS), MTD, age, ECE and seminal vesicle invasion rate. In bootstrapped data set (n = 200 random sampling) algorithm revealed 70.2% sensitivity (95% confidence interval (CI) 58.8–83.0%), 49.9% specificity (95%CI: 42.0–57.7%), 83.9% negative predictive value (NPV; 95%CI: 76.1–91.4%) and 31.1% positive predictive value (PPV; 95%CI: 19.6–39.7%). When limiting analysis to high-risk patients (GS >7) the algorithm improved its performance: sensitivity 91%, specificity 47%, PPV 53%, NPV 89%.

Conclusions
Analyzed algorithm is useful for identifying prostate lobes without ECE and deciding on ipsilateral nerve-sparing technique during radical prostatectomy, especially in patients with GS >7. Due to significant number of false positives in case of: MTD ≥15 mm OR cancer in biopsy ≥15% OR PSA ≥20 ng/mL additional evaluation is necessary to aid decision-making.

Key Words: magnetic resonance ›› clinical prediction rule ›› nomograms ›› prostate specific antigen ›› biopsy ›› PIRADS

INTRODUCTION
Radical prostatectomy constitutes mainstay curative treatment in men with organ-confined prostate cancer [1] as well as in selected patients with locally-advanced disease [2]. Functional outcomes of the surgery, including postoperative continence and potency, might be substantially improved utilizing nerve-sparing [3, 4]. However this approach should be limited to patients in whom tumor does not extend through prostate capsule [5]. Extracapsular extension (ECE) is an independent risk factor of positive surgical margins and might affect long-term biochemical status [4]. Increased understanding
of the effect of ECE on oncological outcomes and the increasing feasibility of unilateral nerve-sparing yielded several predictive tools, that can be used preoperatively to avoid overqualification to nerve-sparing in men suffering from locally advanced disease [6, 7, 8]. We have previously described development and internal validation of logistic-regression-based binary algorithm for safe, side-selective initial qualification to nerve sparing in patients staged preoperatively with mpMRI [9]. The aim of this study was to externally validate our predictive model to assess, whether the results may be generalized to a broader population of patients, who undergo radical prostatectomy with an intention to spare neuro-vascular bundles.

**MATERIAL AND METHODS**

**Patients selection and data collection**

Both development and validation study were approved by relevant ethics committees. A development cohort involved 88 patients (176 lobes) [9], whereas validation cohort involved 154 patients (308 lobes) with prostate cancer who underwent laparoscopic radical prostatectomy with preoperative staging using 3T mpMRI in another department. The images were interpreted by a single experienced radiologist specialized in genitourinary tract diagnostics, who described suspicious prostatic lesions using PIRADSv2 system, according to ESUR recommendations [10]. This resulted in 308 records, because every lobe was considered a separate case, according to the methodology used in the development cohort from the original publication [9]. Patients selection in terms of exclusion and inclusion criteria were as previously described for development cohort [9]. The following clinical variables were extracted for the validation cohort: age, serum prostate specific antigen (PSA), prostate volume, PSA density (PSAD). An extent of tumor infiltration was assessed for each prostate lobe separately using the following variables: number of positive cores, percentage of positive cores, percentage of cancer in total biopsy specimen and Gleason score. If no cancer was found in the lobe at biopsy, all biopsy-derived variables for this side were counted as zero. Consequently, for each side of RP specimen, Gleason score, surgical margins status and pathological stage were reported separately. The analysis of prostate specimens was performed according to the International Society of Urological Pathology guidelines (2014) and the TNM classification was used.

Local staging in the validation cohort was performed using Achieva 3.0-T MRI TX (Philips, Amsterdam, The Netherlands) with dual RF transmitter and 32 independent receiving channels with a multichannel phased-array coil. An endorectal coil was not used. As in the development cohort, examinations included T2-weighted imaging, diffusion-weighted imaging (DWI), and dynamic contrast-enhanced imaging (DCE) carried out according to the European Society of Urological Radiology guidelines. The MRI protocol included: axial T2-weighted turbo spin echo sequence, axial diffusion-weighted imaging spin echo sequence with apparent diffusion coefficient map, axial dynamic contrast-enhanced imaging, axial T1-weighted spin echo with selective fat suppression sequence, axial T1-weighted turbo field echo sequence, coronal and sagittal T2-weighted turbo spin echo sequence in all cases and magnetic

![Figure 1. The use of an algorithm in making decisions about the use of the nerve-sparing technique. The negative and positive predictive values calculated for the validation group were used to determine the extracapsular extension risk. 159 of 308 prostate lobes fulfilled criteria for ECE (-) in the entire validation group. For patients with Gleason score >7 in biopsy the values are 89% and 53% respectively.](image)

**nsRP** – nerve-sparing radical prostatectomy; **MTD** – maximum tumor diameter; **mpMRI** – multiparametric magnetic resonance of prostate; **PSA** – prostate specific antigen; **bx** – biopsy; **ECE** – extracapsular extension; **CI** – confidence interval.
resonance spectroscopic imaging in selected cases. In each case mpMRI was performed more than 4 weeks after prostate biopsy. Examinations in the validation cohort were evaluated by a single radiologist, who was not blinded to the clinical characteristics. The Prostate Imaging Reporting And Data System (PIRADS) version 2 was used to assess lesions. Using this system every suspicious lesion identified in the gland was scored from 1 to 5 based on specific radiologic signs found in mpMRI sequences. Subsequently, an overall score, analogously to Likert scale, was given for each lesion. MTD (maximum tumor diameter) was defined as the largest diameter of a lesion defined PIRADS 3-5. Indications for prostate biopsy included: PSA elevation (>4 ng/mL) and/or abnormal DRE. Tru-Cut prostate biopsy (TRUS core-Bx) was performed systemically and guided with transrectal ultrasound. In case of suspicious lesion visible in TRUS or prebiopsy MRI, an additional targeted core was taken. Histopathological evaluation of biopsy cores and postprostatectomy specimen was as described before [11].

Ethical approval and informed consent

All procedures performed during the study were in accordance with the ethical standards of the institutional and national research committee and with the 1964 Helsinki declaration and its later amendments. Ethical board of Medical University of Warsaw has accepted the retrospective design of the study (approval number AKBE/46/17). Informed consent was obtained from all individual participants included in the study.

Statistical analysis

Continuous variables are presented as medians with corresponding interquartile ranges (IQR). Univariate and multivariate logistic regression models were utilized to confirm the variables from primary model as independent predictors of extracapsular extension. Variables were categorized based on cutoffs identified in development cohort [9]. Area under receiver operating characteristic curve (AUC), sensitivity, specificity, positive (PPV) and negative predictive value (NPV) of the algorithm and its components were identified for bootstrapped cohort. Analyses were performed using SAS 9.4 software (SAS Institute, Cary, USA) and R program (version 3.5.3, the R foundation for Statistical Computing, http://www.r-project.org) with foreign, rms, pROC and ResourceSelection packages was used to perform statistical analysis. The threshold for significance was set at p < 0.05.

**RESULTS**

The development cohort [9] included 88 patients (176 lobes) whereas validation cohort included 154 patients (308 lobes). Mean age in the validation cohort was 63.6 (IQR = 8) and mean PSA was 10.7 ng/mL (IQR = 5.8). In a total of 71 (23.1%) and 21 (6.8%) postprostatectomy lobes specimens ECE and SVI were present respectively. Validation cohort included patients with smaller mean prostate volume (39.9 mL vs 45.5 mL, p = 0.003) but was not significantly different from development cohort regarding age (63.6 vs 63.3 years, p = 0.8), PSA 10.7 vs 10.5 ng/mL (p = 0.9), PSA density (0.30 vs 0.25 ng/ml^2, p = 0.16), total cancerous tissue in biopsy cores in each lobe (22.4% vs 18.1%, p = 0.5), MTD in each lobe (8.9 mm vs 10.1 mm, p = 0.19), ECE prevalence (23.1% lobes vs 30.1%, p = 0.1), SVI prevalence (6.8% lobes vs 8.7%, p = 0.47), biopsy Gleason score ≥7 (32.1% lobes vs 34.1%, p = 0.69) or biopsy Gleason score ≥8 prevalence (6.5% lobes vs 8.5%, p = 0.46). In summary, both cohorts did not differ significantly regarding commonly accepted risk factors for ECE.

Preoperative characteristics of validation cohort are summarized in Table 1. Preoperative characteristics of the development cohort can be found in previous manuscript [9].

**Logistic regression**

To confirm primary model components as independent predictors of ECE, logistic regression was implemented. Uni- and multivariate analysis is presented in Table 2.

### Table 1. Baseline preoperative characteristics of validation cohort

| Variable       | Categories | n (%)       |
|----------------|------------|-------------|
| PSA            | ≤10 ng/mL  | 107 (69.9%) |
| PSA            | 10–20 ng/mL | 32 (20.9%)  |
| PSA            | ≥20 ng/mL  | 14 (9.15%)  |
| Grade group    | I          | 133 (57.33%)|  |
| Grade group    | II         | 59 (25.43%) |  |
| ISUP 2014*     | Grade group I | 133 (57.33%)|  |
| ISUP 2014*     | Grade group III | 20 (8.62%) |  |
| ISUP 2014*     | Grade group IV | 10 (4.31%) |  |
| ISUP 2014*     | Grade group V | 10 (4.31%) |  |
| PSA density    | ≤0.15      | 42 (30%)    |
| PSA density    | 0.15–0.30  | 58 (41.43%) |
| PSA density    | >0.30      | 40 (28.6%)  |

ISUP 2014 GG – International Society of Uropathology 2014 Gleason Groups in biopsy; PSA – prostate specific antigen [ng/mL]; PSAD – prostate specific antigen density [ng/mL^2]; *side-specific data
Multivariate model was constructed from previously implemented predictors (PSA >20 ng/mL, total percentage of cancerous tissue in biopsy cores >15% and MTD >15 mm). The effects of the predictors were lower in the validation sample comparing to the development sample. PSA and percentage of biopsy cancer were confirmed as significant independent predictors and MTD as predictor presenting tendency to significance with 1.9 OR of ECE. Finally, we merged samples to calculate more stable estimation of the effects of the predictors for the entire cohort. Coefficients β were smaller than expected in the validation cohort, which suggested that we tried to validate an overfitted model, but overfitting of the development model was excluded in the original study. The final model form a combined dataset (development + validation) revealed good performance. The regression coefficients in the final model are a compromise between the estimates in the development and validation sample (Table 3). For validation of clinical implementation of the model sensitivity, specificity, positive and negative predictive values were calculated for bootstrapped risk-stratified validation cohort (Table 4).

**DISCUSSION**

In the present study we externally validated previously developed side-specific algorithm predicting extracapsular extension in patients who underwent preprostatectomy mpMRI [9]. Validated model utilizes D’Amico definition of PSA high-risk (≥20 ng/mL) and side-selective variables that describe cancer volume: maximum tumor diameter of suspicious lesion measured in mpMRI and total percentage of cancerous tissue in biopsy cores. Based on nonlinear associations detected in development cohort [9], predefined cut-offs were used to binarize the model. We confirmed fairly good performance of the nomogram in the validation cohort from the external center. The idea of introducing binary model was oriented on frequently unclear risk cut-offs of continuous models and clinical utility of simple risk grouping like presented in D’Amico groups [12], that remain crucial preoperative risk assessment despite constant development of more accurate multivariable models. However, the proposed nomogram was a combination of the nonlinear predictors and, as such, this formula may not be easily reproduced to calculate outcome predictions for new patients. This underlines the need of meticulous validation before broader use. Since outcomes of multiparametric MRI software analysis have been proved to correlate with D’Amico scoring [13], supplementing risk groups with objective MRI-derived measure-

### Table 2. Univariate and multivariate logistic regression for prediction of extracapsular extension

| Variable          | Univariate | Multivariate |
|-------------------|------------|--------------|
|                   | OR (95% CI) | p            | OR (95% CI) | p            |
| ISUP 2014         |            |              |            |              |
| GG1               | 1          |              | 1.06 (1.04–1.09) | 0.0001      |
| GG2               | 2.3 (1.2–4.7) | NS           | 1.06 (1.04–1.09) | 0.0001      |
| GG3               | 1.7 (0.6–4.6) | NS           | 3.9 (1.6–9.4) | 0.003       |
| GG4               | 1.8 (0.6–5.0) | NS           | 4.4 (2.1–9.5) | 0.0001      |
| GG5               | 3.9 (1.4–11.0) | NS          |              |              |
| Age               | 0.99 (0.97–1.03) | NS         |              |              |
| Pvol              | 1 (0.99–1.01) | NS           |              |              |
| MTD [mm]          | 1.07 (1.04–1.09) | <0.0001    | 1.07 (1.04–1.09) | <0.0001    |
| MTD ≥15 mm        | 4.1 (2.6–6.5) | <0.0001      | 9.0 (4.0–15.5) | 0.0003      |
| PSA               | 1.06 (1.04–1.09) | 0.0001     | 1.06 (1.04–1.09) | 0.0001     |
| PSA ≥20 ng/mL     | 3.9 (1.6–9.4) | 0.0003       | 1.8 (0.8–3.3) | 0.19        |
| PSAD              | 4.4 (2.1–9.5) | 0.0001       | 4.4 (2.1–9.5) | 0.0001      |
| Total % of cancer in cores | 1.02 (1.02–1.03) | <0.0001 | 1.02 (1.02–1.03) | <0.0001 |
| Total % of cancer in cores ≥15% | 3.8 (2.4–5.9) | <0.0001 | 3.8 (2.4–5.9) | 0.027       |

### Table 3. Effects of predictors and discrimination indexes in development, validation and combined model

| Model               | Development | Validation | Combined |
|---------------------|-------------|------------|----------|
| PSA ≥20 ng/mL       | 14.3        | 3.9        | 5.0      | <0.0001 |
| % of cancer in cores ≥15% | 7.5        | 2.0        | 3.1      | <0.0001 |
| MTD ≥15 mm          | 7.0         | 1.6        | 3.1      | <0.0001 |

**Discrimination indexes**

|                      | Development | Validation | Combined |
|----------------------|-------------|------------|----------|
| C*                   | 0.858       | 0.641      | 0.744    |
| R^2*                 | 0.499       | 0.098      | 0.241    |

C-index – counted from area under the ROC curve; reproduces diagnostic accuracy [0-1]. R^2* – provides a measure of the proportion of the variance; the larger it is, the more the variance of the dependent variable is explained by the regression model [0-1]; OR – odds ratio; MTD – maximum tumor diameter; PSA – prostate specific antigen.
invented by Partin et al. has become integral part of preprostatectomy track and has been consequent-
ly updated among with stage migration due to PSA screening and mpMRI implementation [6]. The main
limitation of Partins straightforward, extensively validated risk calculator is lack of side-specificity.
Considering extended lymphadenectomy side-specificity is of less significance [18]. To correctly perform
the nerve-sparing surgery a planning method needs a change from ‘patient-approach’ to ‘lobe-approach’.
The side-specific prediction has been proposed in multiple regression-based models. First intro-
buced by Ohori et al included almost all side-specific biopsy volumetric markers as well as grading and
PSA [19] yielding high AUC of 0.773 in external validation [20]. Recently, Sayyid et al have proposed
a similar novel nomogram based on clinical side-
specific variables [7]. External validation of the tool yielded high AUC of 0.74 and excellent calibration.
However, this model is based on multiple subjective variables (TRUS lesion, DRE staging) and neglects
preoperative mpMRI assessment, therefore its clinical implementation might differ depending on depart-
ment. MRI as independent adjunct to clinical data has been introduced by Feng et al. [21], by supple-
menting previous Ohori model with staging mpMRI. Sensitivity and negative predictive value of the up-
dated model have reached 85% and 95% respectively, whereas AUC increased from 0.86 to 0.94. Note-
worthy, validation probe was modest (112 patients), so this encouraging results should be interpreted
with caution due to potential calibration issues. Recently, Alessi et al. released a nomogram based
on the scales proposed by ESUR in the interpreta-
tion of MRI [22]. They have proven, that PIRADS v2 assessment categories of 3 or less rule out the pres-
ence of ECE with great certainty among all PCa-risk groups. However, due to markedly lower specificity of
ESUR scoring systems among intermediate to high PCa-risk group it can be assumed that many of these
men would be incorrectly disqualified from NVB-
sparing surgery, had to rely only on MRI parameters.
Such nomograms also have a serious limitation, as they do not assess a specific risk of ECE for each
side of the gland. The results obtained in this way, however encouraging, may not translate into clinical
practice. An update of the MSKCC nomogram with MRI data provided little, if any, incremental value
risk assessment of ECE improving AUC by only

### Table 4. Sensitivity, specificity, PPV and NPV (bootstrap n = 200)

|                   | Sensitivity (95% CI) | Specificity (95% CI) | PPV (95% CI) | NPV (95% CI) |
|-------------------|----------------------|----------------------|--------------|--------------|
| Development cohort| 91% (83–92)          | 74% (65–98)          | 54% (44–65)  | 94% (89–99)  |
| Validation cohort* (n = 308) | 70% (58–83) | 50% (42–58) | 31% (20–40) | 84% (76–91) |
| Validation cohort (GS = 6)** (n = 152) | 60 (40–76) | 51 (44–59) | 19 (13–27) | 86 (78–93) |
| Validation cohort (biopsy GS = 7)*** (n = 126) | 71 (61–81) | 49 (42–56) | 38 (31–45) | 79 (71–86) |
| Validation cohort (biopsy GS >7)**** (n = 30) | 91 (82–97) | 47 (37–59) | 53 (41–63) | 89 (80–96) |

**Area under curve:**

* 0.65
** 0.60
*** 0.62
**** 0.81

In contrast to majority of previously continuous mod-
el (6–8,18–21), the tool we validated in present study

generates direct decision instead of percentage risk as
outcome. Obviously, such approach further reduces
role of physician judgement. On the other hand sup-
porting decision with the algorithm does not exclude
supplemental use of more accurate tool like nomo-
grams in second step. In fact, by risk grouping based
on volumetric derivates and PSA we intended to de-
termine situation in which extracapsular extension
can be safely ruled out with variables that remain
objective regardless of local experience and preopera-
tive preparation policy. Volume of cancerous tissue
in biopsy seems to meet this condition at first place.
Since mpMRI determined maximum tumor diameter
seems to be more objective and less experience depen-
dent than complete staging assessment and simulta-
neously should be included in every PIRADS report
[11], we believe this variable can be easily used even
in departments with limited experience with MRI.
Understanding that NPV and accuracy of the tool
might be comparable with outcomes of sole staging
mpMRI achieved in highly experienced centers [13,
22, 23], we believe utility of the tool might be mostly
visible when gaining experience with resonance imag-
ing of the prostate. Based on initial studies, MRI was
considered to have no incremental value over stan-
dard staging approach in low-risk patients and was
not indicated in that setting. However, recent meta-
analysis revealed that, in spite of its low sensitivity in
detecting EPE in the low-risk PCa, it provides valu-
able information about feasibility of nerve-sparing
surgery [16]. Moreover, current EAU guidelines rec-
ommend to use prebiopsy MRI for staging purposes.
To provide genuine validation and define best work-
ing environment for the tool, an external cohort of
the department with expertise in prostate can-
treatment was used. Both cohorts were similar
in terms of endpoint prevalence and relevant predic-
tors widely considered to be related to ECE. To sus-
tain certainty that proposed categorical variables can predict ECE independently, logistic regression was performed in validation cohort. We confirmed that total percentage of cancerous tissue in biopsy >15%, MTD >15 mm and PSA >20 ng/mL are ECE predictors in a model validated in Gleason score-stratified external cohort. Given that definitions of predictors were exactly the same and disease advancement was similar, we manage to obtain a fairly good performance at external validation. The discriminative ability dropped at development with c statistic from around 0.86 to 0.64, although Somers’ Dxy rank correlation indicated, that the model still improved the prediction of ECE in the validation cohort by nearly 30% (Table 3). Worse outcomes at validation may be explained by relatively small size of the development cohort, or that patients were originally selected from a single center. Presumably, validation in several external sites would create an opportunity to update this simple formula and enable its border application. The final model was calculated from a combined dataset (development and validation) and revealed good performance (Table 3), similarly to the native model. The calculation of final model derived from a full sample led to more stable estimation of the effect of the predictors and prevented waste of relevant information. This combination of data assumes that the two samples represent a similar population, which is indeed the case (Table 1). Although model managed to safely rule out ECE in entire cohort (NPV 84%), it tended to overestimate the risk, which resulted in high rate of false positives and unsatisfactory specificity as well as PPV. Due to substantial deterioration of sensitivity especially in low risk patients area under ROC curve has not exceeded 0.7 in entire cohort. It is however worth noting, that in patients with high-risk Gleason grade in biopsy (>7), better discrimination resulted not only in further improvement of sensitivity and NPV, but also substantial reduction of false positives (specificity and PPV reaching 60%), which elevated AUC to 0.8. This highlights the impact of different ECE prevalence among distinct PCa-risk groups on diagnostic accuracy. Since positive surgical margins are of major concern in high-risk patients [24] and do not always require change of postoperative track in low risk patients [12] we interpret this results as primary validation outcome that supports its clinical use. The validation results suggest that the algorithm lacks specificity and might additionally underestimate risk in Gleason score ≤7. Thus it cannot be recommended for routine use as a sole preoperative tool assessing ECE risk in every patient. In low- or intermediate-risk patients the algorithm should be used for initial assessment and supplemented with one of the validated nomograms [7, 20, 21], especially in case of men highly motivated to nerve-sparing and unclear contraindications to this approach. In patients with high-risk organ-confined prostate cancer the algorithm can be strongly recommended at least as initial screening and adjunct to nomograms, since its predictive yield seems clearly proven in this group (Figure 1).

The study has several limitations that should be signalized. Although cohorts have not differed significantly considering preoperative assessment as well as risk groups composition, there were some less substantial differences that can impact the validation outcomes. In validation cohort magnetic resonance imaging was performed on 3-T system not 1.5-T like in development cohort. Moreover, validating MRI was performed not only for staging purposes, but also for targeting biopsy in some individuals. That might influence percentage of cancerous tissue in biopsy cores and require additional model calibration with possible change of cut-off. However, clinical use of mpMRI is now wider than in development period and majority of departments use it now also in a prebiopsy setting (25), we believe that this inconsistency might in fact allow validation cohort to reproduce current clinical environment even better. Consequently, validation cohort was collected after ISUP 2014 whereas model was primary developed with ISUP 2005 grading. Surprisingly, validation and development regression models have not found Gleason grade significant any-way. Finally, although ECE prevalence and Gleason composition is similar to those observed in similar series [7, 19, 20, 21], retrospective design might be an obvious source of selection bias. Therefore, authors encourage do perform fully independent validation by other investigators at other centers.

CONCLUSIONS

This external validation confirms the good performance of our model using PSA, biopsy and MRI parameters to predict ipsilateral ECE. Its simplicity and user friendly format provides an easy tool in initial screening of men undergoing radical prostatectomy with the intention of preserving NVB.

CONFLICTS OF INTEREST
The authors declare no conflicts of interest.

AVAILABILITY OF DATA AND MATERIALS
The datasets used and analyzed during the current study are available from the corresponding author on reasonable request.

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AUTHORS’ CONTRIBUTION
All authors listed on the manuscript have contributed significantly to the study. All authors have been involved in the writing of the manuscript at draft and any revision stages, and contributed to the final version of manuscript. All authors have read and approved the final version.

References

1. Bill-Axelson A, Holmberg L, Garmo H, et al. Radical prostatectomy or watchful waiting in early prostate cancer. N Engl J Med. 2014; 370: 932-942.
2. Sheng W, Zhang H, Lu Y. Survival outcomes of locally advanced prostate cancer in patients aged <50 years after local therapy in the contemporary US population. Int Urol Nephrol. 2018; 50: 1435-1444.
3. Steineck G, Bjartell A, Hugosson J, et al. Degree of preservation of the neurovascular bundle during radical prostatectomy and urinary continence 1 year after surgery. Eur Urol. 2015; 67: 559-568.
4. Salonia A, Burnett AL, Graefen M, et al. Prevention and management of postprostatectomy sexual dysfunctions part 2: recovery and preservation of erectile function, sexual desire, andorgasmic function. Eur Urol. 2012; 62: 273-286.
5. Ward JF, Zincke H, Bergstralh EJ, Slezak JM, Myers RP, Blute ML. The impact of surgical approach (nerve bundle preservation versus wide local excision) on surgical margins and biochemical recurrence following radical prostatectomy. J Urol. 2004; 172 (Pt 1): 1328-1332.
6. Tosoi D, Chappidi M, Feng Z, et al. Prediction of pathological stage based on clinical stage, serum prostate-specific antigen, and biopsy Gleason score: Partin Tables in the contemporary era. BJU Int. 2017; 119: 676-683.
7. Sayyid R, Perls N, Ahmad A, et al. Development and external validation of a biopsy-derived nomogram to predict risk of ipsilateral extraprostatic extension. BJU Int. 2017; 120: 76-82.
8. Zorn KC, Gallina A, Hutterer GC, et al. External validation of a nomogram for prediction of side-specific extracapsular extension at robotic radical prostatectomy. J Endourol. 2007; 21: 1345-1351.
9. Zapala P, Dybowski B, Bres-Niewada E, et al. Predicting side-specific prostate cancer extra-capsular extension: a simple decision rule of PSA, biopsy, and MRI parameters. Int Urol Nephrol. 2019; 51: 1545-1552.
10. Weinreb JC, Barentsz JO, Choyke PL, et al. PI-RADS Prostate Imaging-Reporting and Data System: 2015, Version 2. Eur Urol. 2016; 69: 16-40.
11. Kozikowski M, Zagóźdzon B, Gola M, Dobruch J. Prostate Imaging Reporting and Data System in prostate cancer staging and planning for radical prostatectomy. Videocirurgia Inne Tech Maloinwazyjnej Video surgery Miniinvasive Tech. 2019; 14: 262-270.
12. Boorjian SA, Kamm J, Rangel LI, Bergstrah EJ, Blute ML. Mayo Clinic validation of the D’Amico risk group classification for predicting survival following radical prostatectomy. J Urol. 2008; 179: 1354-1360.
13. Hameed M, Ganeshan B, Shur J, Mukherjee S, Afaq A, Batura D. The clinical utility of prostate cancer heterogeneity using texture analysis of multiparametric MRI. Int Urol Nephrol. 2019; 51: 817-824.
14. Barentsz JO, Weinreb JC, Verma S, et al. Synopsis of the PI-RADS v2 Guidelines for Multiparametric Prostate Magnetic Resonance Imaging and Recommendations for Use. Eur Urol. 2016; 69: 41-49.
15. Kumar A, Gupta P, Kumar S, et al. 3-D transperitoneal laparoscopic radical prostatectomy in locally advanced high-risk prostate cancer: a prospective evaluation. Cent Eur J Urol. 2019; 72: 218-219.
16. Kozikowski M, Malewski W, Michalak W, Dobruch J. Clinical utility of MRI in the decision-making process before radical prostatectomy: Systematic review and meta-analysis. PloS One. 2019; 14: e0210194.
17. Ross P, Gerigk C, Gonen M, et al. Comparisons of nomograms and urologists’ predictions in prostate cancer. Semin Urol Oncol. 2002; 20: 82-88.
18. Milonas D, Venclovaz Z, Muijik T, Jievaltas M, Joniau S. External validation of Memorial Sloan Kettering Cancer Center nomogram and prediction of optimal candidate for lymph node dissection in clinically localized prostate cancer. Cent Eur J Urol. 2020; 73: 19-25.
19. Ohori M, Kattan MW, Koh H, et al. Predicting the presence and side of extracapsular extension: a nomogram for staging prostate cancer. J Urol. 2004; 171: 1844-1849.
20. Satake N, Ohori M, Yu C, et al. Development and internal validation of a nomogram predicting extracapsular extension in radical prostatectomy specimens. Int J Urol Off J Jpn Urol Assoc. 2010; 17: 267-272.
21. Feng TS, Sharif-Afshar AR, Wu J, et al. Multiparametric MRI Improves Accuracy of Clinical Nomograms for Predicting Extracapsular Extension of Prostate Cancer. Urology. 2015; 86: 332-337.
22. Alessi S, Pricolo P, Summers P, et al. Low PI-RADS assessment category excludes extraprostatic extension (≥pT3a) of prostate cancer: a histology-validated study including 301 operated patients. Eur Radiol. 2019; 29: 5478-5487.
23. Weaver JK, Kim EH, Vetter JM, et al. Prostate Magnetic Resonance Imaging Provides Limited Incremental Value Over the Memorial Sloan Kettering Cancer Center Preradical Prostatectomy Nomogram. Urology. 2018; 113: 119-128.
24. Brimo F, Partin AW, Epstein JI. Tumor grade at margins of resection in radical prostatectomy specimens is an independent predictor of prognosis. Urology. 2010; 76: 1206-1209.
25. Baboujadian M, Bandelier Q, et al. MRI-targeted biopsy for detecting prostate cancer: have the guidelines changed our practices and our prostate cancer detection rate? Int Urol Nephrol. 2020; 52: 611-618.