Evaluation of the contribution of tenderness, juiciness, and flavor to the overall consumer beef eating experience

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ABSTRACT: The objectives of this study were to evaluate the contribution of tenderness, juiciness, and flavor to the overall consumer beef eating experience and to evaluate the risk of overall palatability failure due to the unacceptable level of one or more of these traits. Data from 11 previously conducted studies representing a wide range of treatments and levels of eating quality that included more than 1,500 beef samples and 1,800 consumers were compiled and analyzed for this study. Results of a multivariate regression indicated that tenderness, flavor, and juiciness accounted for 43.4%, 49.4%, and 7.4%, respectively, of overall palatability ($P < 0.05; R^2 > 0.99$). Additionally, the odds of a steak being rated unacceptable overall when tenderness, juiciness, or flavor were rated unacceptable were 2.2 to 1 (69%), 1.9 to 1 (66%), and 3.3 to 1 (77%), respectively. This indicated overall palatability was 7.2, 6.5, and 12.3 times more likely to be rated unacceptable if tenderness, juiciness, or flavor, respectively, was also rated unacceptable. Additionally, the percentage of samples rated acceptable for each palatability trait increased ($P < 0.05$) as quality grade increased. More than 88% of USDA Prime samples were rated acceptable for each palatability trait, whereas only 74.8–77.3% of USDA Select samples were rated acceptable for each palatability trait. Marbling score accounted for 14–16% of the variation ($P < 0.01$) in consumer palatability scores for each trait and intramuscular fat percentage accounted for 17–21% of the variation in each trait ($P < 0.01$). Logistic equation models for the predicted probability of an acceptable rating for each palatability trait based on intramuscular fat percentage accounted for only a minimal amount of variation ($P < 0.01; R^2 \leq 0.09$). Results of this study indicate the relative contribution of tenderness, juiciness, and flavor to overall palatability. They provide evidence that the failure of even a single palatability trait dramatically increases the likelihood of overall palatability failure, indicating that no single palatability trait is most important, as beef palatability is dependent upon the acceptance of all three traits: tenderness, juiciness, and flavor.

Key words: beef, flavor, juiciness, marbling, palatability, tenderness

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Transl. Anim. Sci. 2018.2:26–36
doi: 10.1093/tas/txx008

INTRODUCTION

Overall beef palatability can be attributed to three primary traits—tenderness, juiciness, and flavor (Smith and Carpenter, 1974). It has been widely documented that consumers are willing to pay premiums for beef that will repeatedly...
and consistently meet their eating expectations (Shackelford et al., 2001; Lyford et al., 2010). Countless studies over the past eight decades have evaluated the impact of various animal production factors (breed, genetics, diet, animal health, growth-promotant use, animal age, handling stress, etc.) and meat production and processing factors (USDA quality grade, marbling, aging, electrical stimulation, chilling methods, carcass suspension method, product enhancement, etc.) on these traits. As a result, a greater understanding is known today about the factors impacting beef eating quality than ever before. Much of this research has resulted in industry-wide changes to production practices, as well as the development of new technologies aimed at improving beef eating quality and consumer satisfaction.

Multiple authors have worked to identify which of these palatability traits contributes the most to overall eating satisfaction. These studies have historically identified tenderness as the most important palatability trait (Savell et al., 1987; Miller et al., 1995a; Savell et al., 1999; Egan et al., 2001). Results of the first three National Beef Quality Audits in 1991, 1995, and 2000 identified beef tenderness as one of the most important quality challenges facing the beef industry (Smith et al., 1992; Boleman et al., 1998; McKenna et al., 2002). As a result, much industry research over the past 25 years has focused on tenderness improvement. This focus has resulted in large improvements in tenderness as indicated by the results of the most recent Beef Tenderness Survey, which showed that over 95% of beef from the rib and loin at the retail level would be classified as “tender” or “very tender” (Savell et al., 2016). As a result of these improvements in beef tenderness, more recent investigations have most commonly shown flavor to be the largest factor impacting overall beef eating satisfaction (Killinger et al., 2004; O’Quinn et al., 2012; Corbin et al., 2015; Lucherk et al., 2016).

Regardless, such inferences about a single palatability trait’s impact on beef eating quality are oversimplified. Overall beef eating quality is dependent upon all three factors—tenderness, juiciness, and flavor—as well as the interaction among these traits. Beef steaks may excel at one or even two of these traits, yet fail to meet consumer eating expectations due to the unsatisfactory level of another trait. Conversely, a steak may be deemed acceptable by consumers due primarily to the outstanding level of a single trait despite the lower or even unacceptable levels of one or both of the other traits. To date, no comprehensive study has evaluated this interaction among palatability traits and assessed the relative risk of an unacceptable overall eating experience associated with the failure of a single or combination of palatability traits. It was therefore the objective of this study to evaluate the relative contribution of tenderness, juiciness, and flavor to overall consumer eating satisfaction.

**MATERIALS AND METHODS**

Over the past 6 yr, a number of consumer studies (Table 1) have been conducted evaluating beef palatability (O’Quinn et al., 2012; Hunt et al., 2014; Cashman, 2015; Corbin et al., 2015; Legako et al., 2015; Lucherk et al., 2016; Ron et al., 2016; Wilfong et al., 2016; Gredell et al., 2015; McKillip et al., 2017; Vierck, 2017). The beef samples in these studies were from a wide range of treatments representing a large diversity of beef eating quality levels. Most of these studies have included a diversity of USDA quality grades, as well as differences in muscle, degree of doneness, enhancement level, finishing diet, and animal maturity. Together, this group of samples represented a large variation in beef eating quality and interactions among tenderness, juiciness, and flavor acceptability.

Within each study, the same 100-mm line scales were used for consumer evaluation of steak tenderness, juiciness, flavor, and overall liking. Scales were anchored as extremely tough/dry/dislike extremely at the 0 end point and extremely tender/juicy/like extremely at the 100 end point. Additionally, consumers rated each trait as either acceptable or unacceptable (yes/no), providing definitive consumer perceptions of steak acceptability for each trait.

All samples used in these studies were cooked using similar dry-heat grilling procedures. Collectively, these studies included more than 1,800 beef consumers from multiple regions of the United States and included 1,505 unique samples resulting in more than 12,000 individual consumer observations. This large number of consumer observations utilizing the same scaling and similar testing procedures represented a unique opportunity for a robust analysis evaluating the contribution of the three most important palatability factors—tenderness, juiciness, and flavor—to consumer overall eating experience. Moreover, most of the samples in the dataset included the USDA marbling score as well as the percentage of intramuscular fat. Because of this, the contribution of marbling and fat percentage to each palatability trait and overall eating experience was also evaluated.

Many of the studies used in this data set included differences in USDA quality grade and its effect on
palatability. Because of this, the data offered an opportunity to analyze and evaluate the contribution of marbling level and intramuscular fat content to beef eating quality. Numerous authors have previously evaluated this relationship; however, this set of data was restricted to only consumer panelists and therefore offered the opportunity to gain additional insight into the impact of marbling on consumer sensory perceptions. For these analyses, the data were restricted to only steaks from the longissimus dorsi that were from young (A maturity) grain-finished cattle that were cooked to a medium (71 °C) degree of doneness. Additionally, samples that had been enhanced with moisture solutions were excluded from these analyses.

### Statistical Analysis

To accomplish our objectives, the raw data from all studies were compiled as a single dataset. The average sensory score for each palatability trait was determined for each sample by averaging across the individual consumer ratings for the sample. All statistical analyses were conducted using SAS (SAS Version 9.4, Cary, NC). A multivariate regression model was constructed using the sample means to determine the relative contribution of tenderness, juiciness, and flavor to consumer overall liking scores using PROC REG and the stepwise selection procedure, with variables required to be significant \( P < 0.05 \) to enter the model and to remain in the final model. The odds ratios and relative risk of an unacceptable overall eating experience were determined based on the acceptability of the three individual sensory traits and calculated using PROC GENMOD. Moreover, the percentage of samples from each quality grade that were identified as acceptable for each trait were determined and analyzed using PROC GLIMMIX with a model that included a binomial error distribution. Lastly, simple linear regression models were calculated using PROC REG to quantify the contribution of intramuscular fat percentage to consumer sensory ratings, and logistic regression models were determined using PROC LOGISTIC to estimate the probability of sensory trait acceptance based on intramuscular fat content.

| Study                  | Treatments used                                                                 | Muscles                                         | Number of consumers | Number of samples |
|------------------------|---------------------------------------------------------------------------------|-------------------------------------------------|---------------------|-------------------|
| O’Quinn et al. (2012)  | Prime, High choice\(^1\), Low choice\(^2\), Select, Standard, Wagyu             | Longissimus dorsi                               | 120                 | 72                |
| Hunt et al., 2014      | Top choice\(^1\), Select                                                       | Longissimus dorsi                               | 120                 | 108               |
| Corbin et al. (2015)   | Prime, High choice\(^1\), Low choice\(^2\), Select, Standard, Wyagyu, Grass fed| Longissimus dorsi                               | 120                 | 64                |
| Legako et al. (2015)   | Prime, High choice\(^1\), Low choice\(^2\), Select, Standard                    | Longissimus dorsi                               | 278                 | 80                |
| Cashman (2015)         | Prime, Top choice\(^1\), Low choice\(^2\), Select, Standard; Young and Grain-finished mature | Longissimus dorsi                               | 120                 | 150               |
| Lucherk et al. (2016)  | Prime, Top choice\(^1\), Low choice\(^2\), Select, Standard, Enhanced select   | Longissimus dorsi                               | 252                 | 252               |
| Wilfong et al. (2016)  | Prime, Top choice\(^1\), Low choice\(^2\), Select, Select; angus                | Longissimus dorsi                               | 112                 | 80                |
| Ron et al. (2016)      | Top choice\(^1\), Select, Grass finished                                       | Longissimus dorsi                               | 240                 | 240               |
| Gredell et al. (2015)  | Select, Standard; Young, Grain-finished mature, and Forage-finished mature      | Longissimus dorsi                               | 120                 | 90                |
| McKillip et al. (2017) | Prime, Low choice\(^1\), Low select\(^2\), Enhanced prime, Enhanced low choice, Enhanced select | Longissimus dorsi                               | 252                 | 252               |
| Vierck (2017)          | Top choice\(^1\), Low choice\(^2\), Select                                      | Longissimus dorsi                               | 104                 | 117               |
| Total                  |                                                                                  |                                                 | 1,838               | 1,505             |

\(^1\)High choice: marbling score of Moderate\(^0\)–Moderate\(^1\)  
\(^2\)Low choice: marbling score of Small\(^0\)–Small\(^1\)  
\(^3\)Top choice: marbling score of Modest\(^0\)–Moderate\(^1\)  
\(^4\)Low select: marbling score of Slight\(^0\)–Slight\(^1\)
For this study, the odds of an unacceptable overall palatability rating based on the acceptability of the other palatability traits were calculated. The odds are reported as a ratio of the number of unacceptable ratings to the number of acceptable ratings. In this way, odds of 3 to 1 would represent that for every three unacceptable ratings, there was only one acceptable rating, or there was a 75% chance (3/ [3 + 1]) of an unacceptable overall palatability rating.

RESULTS AND DISCUSSION

Modeling Beef Palatability

To develop a model for overall beef eating satisfaction, a multivariate regression analysis was used. Sample means from all 11 studies were included in the regression analysis to account for and include a vast amount of variation in eating quality associated with the various treatments (muscle, degree of doneness, marbling level, animal diet, etc.). For this analysis, sample overall liking scores were used as the dependent variable, and consumer tenderness, juiciness, and flavor liking scores as well as their interactions were used as explanatory variables. Additionally, the intercept was determined to be highly nonsignificant (P > 0.70) and was therefore excluded from the model.

The final palatability model determined (P < 0.01) was as follows: Consumer overall liking = (0.42 × tenderness) + (0.07 × juiciness) + (0.48 × flavor). This model accounted for greater than 99% of the variation (R² > 0.99) in consumer overall liking scores. This provides evidence that the linear combination of tenderness, juiciness, and flavor accounts for practically all of the variation in overall consumer eating satisfaction. The stepwise procedure used required all variables in the final model to be significant (P < 0.05), indicating that each of these three traits accounted for unique variation in consumer overall liking scores.

It is also noteworthy that the interaction terms among the three traits never entered the model, as they were nonsignificant (P > 0.05). This indicates that the effects of tenderness, juiciness, and flavor on overall eating satisfaction are not dependent upon the level of the other traits. It was hypothesized that the interaction of the traits may provide some synergy (i.e., when all traits were at a certain positive level, the impact on overall eating quality would be greater than the sum; vice versa for negative trait evaluations). However, no such interaction was found in our analysis.

The traits of tenderness, juiciness, and flavor have long been considered the most important palatability traits affecting beef eating quality. Some authors have also considered the impact of other factors such as texture, mouthfeel, appearance, and odor (Smith and Carpenter, 1974; Watson et al., 2008a). Though these factors were not evaluated in this study, our regression results would indicate that though other factors may play a role in the perception of beef palatability, the contribution of these factors is likely also accounted for by measures of tenderness, juiciness, and flavor.

When evaluating the contribution of each trait to overall eating quality, the equation indicates that flavor contributes the most (49.4%), followed by tenderness (43.4%) and juiciness (7.4%). Additionally, changes in tenderness and flavor on overall eating quality are very close to equal, with a 1 unit change in flavor rating corresponding to a 1.1 unit change in tenderness. However, much larger changes in juiciness are required to equal changes in overall palatability compared with the other two traits. A change in juiciness of 5.9 units is needed to equal a 1 unit change in tenderness, and a 6.7 unit change in juiciness is needed to equal a 1 unit change in flavor.

Many studies have asked consumers which palatability trait is the most important when consuming beef. Consumers from two such studies conducted in the mid-90s reported that the majority of consumers (50–51%) indicated tenderness was most important, followed by flavor (39–40%) and the fewest (10%) stating juiciness (Miller et al., 1995b; Huffman et al., 1996). More recent studies have indicated a shift in these percentages, with the majority of today’s consumers indicating flavor (49–55%) as the most important trait, with fewer consumers (36–40%) now identifying tenderness as most important (Chail et al., 2016; Wilfong et al., 2016; Chail et al., 2017; McKillip et al., 2017). However, the percentage of consumers who identify juiciness as most important has remained relatively constant, at close to 10% (Chail et al., 2016; Wilfong et al., 2016; Chail et al., 2017; McKillip et al., 2017). This shift in consumer emphasis away from tenderness and to beef flavor is reflected in the relative contribution of each trait in the determined regression model.

Other authors have reported regression models including the three palatability traits accounted for only 79% of the variation in overall palatability score, with flavor alone accounting for the most variation (67%; Huffman et al., 1996). However, in that study the authors used 8-point hedonic scales.
as opposed to the line scales used in the current studies, perhaps explaining some of the observed difference. Taking a different approach, Platter et al. (2003) modeled the effect of tenderness, juiciness, and flavor on overall sample acceptability (yes/no) as opposed to overall palatability score. In that study, the authors reported that logistic regression equations for each trait alone explained over 50% of the variation in overall acceptability and the three variable model explained 62% of the total variation (Platter et al., 2003).

To date, the most extensive modeling of beef palatability as it relates to tenderness, juiciness, and flavor is the Meat Standards Australia (MSA) grading system. The MSA grading system differs from the U.S. system in many ways. One of the primary differences is that the MSA grading system identifies and grades beef cuts (steaks, roasts, stew meat, stir-frys, etc.) based on predicted eating quality as opposed to carcasses, as is done in the United States (Polkinghorne et al., 2008). Cuts are assigned to one of four quality categories—premium quality (five star), better than everyday quality (four star), good everyday quality (three star), and unsatisfactory (Watson et al., 2008a, 2008b). This is done through the calculation of a composited estimate of the eating quality (MQ4) based on an equation considering a variety of animal production and meat processing inputs (Watson et al., 2008b).

At the heart of the MSA grading system is an ever-growing consumer database that has grown in consumer observations and sample numbers since it was first started in the late 1990s. These consumer observations provided the modeling information needed for the development of the MQ4 score. MQ4, as an estimate of composited eating quality, was originally calculated by the equation: MQ4 = (0.4 × tenderness) + (0.1 × juiciness) + (0.2 × flavor) + (0.3 × overall palatability) (Watson et al., 2008a). In this way, the original MSA model weighted beef eating quality as 40% tenderness, 30% overall palatability, 20% flavor, and 10% juiciness. In more recent years, this model has been updated to reflect the growing impact of flavor, with the newest model weighting tenderness, flavor, and overall palatability equally at 30%, with juiciness weighted at 10% (Legako et al., 2015). A primary difference between the MSA palatability model and the calculated model in our study is the inclusion of overall palatability score as an independent explanatory variable in the MSA model as opposed to the dependent response variable as was done in our study. An additional difference between the two models is in the influence of flavor. In the model in our study, flavor was the largest contributor to overall eating quality; however, in the current MSA model, flavor is equally weighted with both tenderness and overall palatability score. Nevertheless, the MSA model provides another indication of the relative contribution of tenderness, juiciness, and flavor to overall beef eating quality.

The palatability model determined for our study reflects the observed emphasis placed on beef flavor by today’s beef consumers. A greater percentage of consumers have self-reported the importance of flavor in recent years, and the model indicates that this emphasis is reflected in their beef eating experiences. Tenderness, however, remains the second largest driver of overall eating quality, with a large portion of the beef eating experience dependent upon tenderness ratings. As the beef industry continues to change over the next few decades, and with a recent industry focus on beef flavor improvement, this palatability model should be reevaluated at regular intervals (every 5 or 10 yr) to monitor the impact of industry changes impacting flavor, tenderness, and juiciness and their relative contribution to consumer eating satisfaction.

### Odds of Overall Palatability Failure

When evaluating the contribution of tenderness, juiciness, and flavor to overall eating quality, it is important to determine the relative risk of a product failing overall due to the failure of one or more of the specific palatability traits. Throughout all of the included studies, yes/no acceptability questions were asked for the each of the traits, allowing the consumers to make a definitive assessment of whether or not the sample met their expectations for that trait. Table 2 provides the estimates for the likelihood of overall failure based on the failure/acceptance of the other traits.

Odds ratios represent the relative increase in the odds of an event occurring (overall palatability failing) due to another event (unacceptable rating for tenderness, juiciness, or flavor). For example, in Table 2, the odds of overall palatability failing when tenderness is acceptable is 1 in 10 (10% chance), whereas the odds of overall palatability failing when tenderness is unacceptable is 2.2 to 1 (69% chance). Therefore, the odds ratio is 20.8 (odds when tenderness is unacceptable/odds when tenderness is acceptable). So the odds of overall palatability failing when tenderness is unacceptable are 20.8 times higher than when tenderness is acceptable. The relative risk is the increased risk of an event occurring (overall unacceptable) due to 20.8 times higher than when tenderness is acceptable.
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another event (unacceptable tenderness). Thus, the likelihood of unacceptable overall palatability is 7.2 times higher when tenderness is unacceptable.

With respect to flavor, only 1 in 15 (6.7% chance) steaks failed for overall palatability when flavor was also acceptable; however, this increases to 3.3 to 1 (76% chance) when flavor was unacceptable. The odds of overall palatability failing when flavor was unacceptable are 49 times higher than when flavor was acceptable, and overall palatability failure is 12.3 times more likely due to unacceptable flavor.

For juiciness, one in every nine steaks (11% chance) were unacceptable overall when juiciness was acceptable compared with close to two out of every three (66% chance) when juiciness was unacceptable. This indicates that overall palatability was 6.5 times more likely to fail when juiciness was unacceptable, with the odds of failure 17.1 times greater due to juiciness failure. Though juiciness contributes only 10% to overall palatability, these results indicate the large impact that even juiciness failing to meet consumer expectations can have on overall palatability.

When more than one palatability trait failed, the odds of overall palatability failure increased dramatically. Most notably, when tenderness and flavor were both unacceptable, the odds of overall palatability failing were 516.5 times greater than when both traits were acceptable, with overall palatability more than 46 times more likely to fail when both traits were unacceptable. When juiciness failed in combination with tenderness or flavor, the odds of overall palatability failure were increased 92 or 294 times, respectively. Lastly, when all three traits were acceptable, only 1 in every 93 steaks (~1% chance) were unacceptable overall. However, when all three traits failed, the odds of failure increased almost 2,000 times to more than a 95% chance, and the likelihood of overall failure was 89.5 times more likely.

These results indicate the significant consequences to overall beef palatability if one or more of the individual palatability traits are viewed as unacceptable by consumers. Though the regression analysis in the previous section provided insight to the relative contribution of each of these traits to overall eating quality, results presented in Table 2 demonstrate the importance of assuring that all three traits are at an acceptable level. Efforts to improve beef palatability focused on only a single trait should ensure that improvements in the single area are not at the detriment of the other two. The odds of beef failing to meet a consumer’s overall eating expectations are increased significantly if even just one of the three individual traits is unacceptable.

Impact of Marbling and Intramuscular Fat Level
The percentage of samples from each quality grade that were rated as acceptable from the yes/no acceptability questions for each of the palatability traits are presented in Table 3. The percentage of acceptable samples increased ($P < 0.01$) for tenderness (Prime > Premium choice = Low choice > Select = Standard), juiciness (Prime > Premium choice = Low choice > Select = Standard), flavor (Prime > Low choice > Select = Standard), and overall palatability (Prime > Premium choice > Low choice > Select = Standard) as quality grade increased.

Table 2. Odds of an unacceptable eating experience based on tenderness, juiciness, and flavor acceptability

| Palatability trait | Odds when trait is acceptable | Odds when trait is unacceptable | Odds ratio | Relative risk |
|--------------------|-------------------------------|--------------------------------|------------|--------------|
| Tenderness         | 1 in 10                       | 2.2 to 1                       | 20.8       | 7.2          |
| Juiciness          | 1 in 9                        | 1.9 to 1                       | 17.1       | 6.5          |
| Flavor             | 1 in 15                       | 3.3 to 1                       | 49.0       | 12.3         |
| Tenderness and juiciness | 1 in 15               | 6.3 to 1                       | 92.0       | 13.5         |
| Tenderness and flavor | 1 in 50                     | 10.3 to 1                      | 516.5      | 46.8         |
| Juiciness and flavor | 1 in 35                     | 8.3 to 1                       | 293.7      | 32.4         |
| Tenderness, juiciness, and flavor | 1 in 93               | 21.5 to 1                      | 1989.1     | 89.5         |

1Odds of overall eating experience failing when individual palatability trait is rated acceptable. Odds reported as number of unacceptable ratings in given number of observations.
2Odds of overall eating experience failing when individual palatability trait is rated unacceptable. Odds reported as number of unacceptable ratings to number of acceptable ratings.
3Relative increase in odds of unacceptable eating experience when trait is rated unacceptable (i.e., Odds of failure are X times greater than when trait is acceptable).
4Increased risk of unacceptable eating experience when trait is unacceptable (i.e., Overall unacceptable rating is X times more likely than when trait is acceptable).
Table 3. Percentage of grain-finished strip loin steaks of various USDA quality grades cooked to a medium degree of doneness rated as acceptable by consumers

| USDA quality grade | Tenderness | Juiciness | Flavor | Overall liking |
|--------------------|------------|-----------|--------|---------------|
| Prime              | 95.14\textsuperscript{a} | 92.42\textsuperscript{a} | 88.11\textsuperscript{a} | 91.37\textsuperscript{a} |
| Premium choice\textsuperscript{1} | 86.61\textsuperscript{b} | 84.97\textsuperscript{b} | 85.44\textsuperscript{b} | 86.83\textsuperscript{b} |
| Low choice\textsuperscript{2} | 86.31\textsuperscript{a} | 83.33\textsuperscript{b} | 83.83\textsuperscript{b} | 83.08\textsuperscript{b} |
| Select             | 77.30\textsuperscript{b} | 75.96\textsuperscript{b} | 75.38\textsuperscript{b} | 74.75\textsuperscript{b} |
| Standard           | 74.53\textsuperscript{b} | 67.99\textsuperscript{c} | 72.29\textsuperscript{c} | 72.04\textsuperscript{c} |
| SEM\textsuperscript{3} | 1.81      | 1.94      | 1.86   | 1.86          |
| \textit{P} value  | <0.01     | <0.01     | <0.01  | <0.01         |

\textsuperscript{1}Premium choice: marbling score of Modest\textsuperscript{5}–Moderate\textsuperscript{5}.

\textsuperscript{2}Low choice: marbling score of Small\textsuperscript{5}–Small\textsuperscript{5}.

\textsuperscript{3}SEM (largest) of the least squares means.

\textsuperscript{abcd}Means in the same column lacking a common superscript differ (\textit{P} \textless 0.05).

These results indicate that USDA quality grade effectively sorts steaks based upon palatability trait acceptability, with higher USDA quality grades having a higher percentage of steaks rated acceptable for each trait than lower grades. Prime had more than 91% of samples rated acceptable for all traits other than flavor, representing a greater (\textit{P} \textless 0.05) percentage than all lower grading beef. Conversely, almost 25% of Select beef failed to meet consumer expectations for all palatability traits and had a similar (\textit{P} \textgt 0.05) percentage of samples rated unacceptable for all traits, other than flavor, as Standard. Currently, close to 20% of cattle nation-wide grade USDA Select (USDA, 2017b). Taken together, this represents a large challenge for the U.S. beef industry. With such a large portion of lower grading (\textless Select) product failing to meet consumer eating expectations combined with the high percentage of carcasses currently in these grades, this represents a significant amount of beef product that will ultimately fail to meet consumer eating expectations. Moreover, the data presented in Table 3 included only strip steaks cooked to medium. The actual observed failure rate for Select beef is likely much higher considering the proportion of consumers who cook steaks to greater than medium (Reicks et al., 2011) and the negative impact of increased degree of doneness on beef palatability. The study by Lucherk et al. (2016) reported Select and Standard steaks cooked to well done were rated unacceptable overall 46% and 50% of the time, respectively. In that study, only consumers who preferred well-done evaluated the well-done samples. There is the potential that this number could be much higher for consumers who prefer lower degrees of doneness and mistakenly overcook steaks at home or are served an overcooked steak at foodservice. Additionally, the failure rate for many of the tougher muscles in the carcass that are traditionally cooked via dry-heat methods from Select may fail at a much higher rate as well.

It is also interesting to note that Premium choice (upper 2/3 of Choice grade) had a greater portion of samples rated acceptable overall than Low choice; however, a similar percentage of samples rated acceptable for each palatability trait. This advantage in overall palatability and demand by consumers is reflected in the premiums garnered by the wholesale cut prices of this category over commodity Choice products (USDA, 2017a).

These results differ from previous authors who have evaluated the probability of an unsatisfactory eating experience based on quality grade. A study by Smith et al. (2008) compiled results from 14 previous works and determined the probability of an unsatisfactory eating experience for Prime to be 1 in 33 (3%), Premium choice to be 1 in 10 (10%), Low choice to be 1 in 6 (16%), Select to be 1 in 4 (25%), and Standard to be 1 in 2 (50%). Additionally, a more recent report by Tatum (2015) composited results from five more recent studies and reported the odds of failure to be 1 in 33.6, 1 in 13.8, 1 in 5.4, 1 in 2.9, and 1 in 2.2 for Prime, Premium choice, Low choice, Select, and Standard, respectively. Results from this study indicate that the odds of Prime steaks failing to produce an acceptable eating experience are actually much higher (1 in 10.6) than reported by these previous authors. Moreover, the probability of Premium choice producing an unacceptable eating experience is also higher (1 in 6.6) than previously reported. Estimates for the failure rate in the current analysis were similar to those reported by Tatum (2015) for Low choice (1 in 4.9 vs. 1 in 5.4); however, this study indicated that the probability of a satisfactory eating experience in the lower to grades was actually higher than the values reported by Tatum (2015): 75% vs. 66% for Select and 72% vs. 55% for Standard.
The observed differences between the current work and that of Smith et al. (2008) and Tatum (2015) are likely the result of the differences in study types used for the analyses. Both Smith et al. (2008) and Tatum (2015) included studies in their analyses that comprised trained sensory panelists. Trained panels are designed to evaluate sensory traits as objectively as possible. Additionally, trained sensory panelists must complete a training or orientation procedure to insure that the panelists are assessing all traits similarly, and that the amount of variation among samples as assessed across the panel is minimal. Because of this, the data from trained sensory panelists should not be interpreted the same as results from consumer panelists who assess samples based on their own individual biases and interpretations. Trained panelists are also much more discriminating than consumer panelists, allowing for greater separation among samples of different treatments and quality grades. Additionally, many of the works used by these authors did not contain yes/no acceptability questions, leaving the authors to have to make a judgment based on the sensory score on the rating scale as to whether or not the sensory panelist would have considered the overall eating experience satisfactory (acceptable).

These differences in the types of studies used and the method for determination of negative/unacceptable eating experience are likely responsible for the observed differences. It appears the use of the trained panel data in these analyses skewed the likelihood of negative eating experiences for the highest marbled (Prime and Premium choice) and the lowest marbled (Select and Standard) grades. The previous reports underestimated the likelihood of product unacceptability for the high marbled products and overestimated the likelihood in the low marbled products. The data presented in the current report are reflective of the percentage of samples from these quality grades that were rated acceptable by consumers (yes/no) and provide one of the best estimates of the actual failure rate for these grades as consumed by beef end users and consumers.

Simple linear regressions were also performed to assess the impact of marbling level and fat percentage on each of the palatability traits (Table 4). Marbling content was related ($P < 0.01$) to all four palatability traits, but explained only 15%, 16%, 14%, and 16% of the variation in sensory panel tenderness, juiciness, flavor, and overall liking ratings, respectively. Differences in personnel assessing marbling level to the 10th of a marbling score and the high level of variability among consumer panelists make modeling the relationship difficult, especially across the number of studies used within this analysis. The percentage of intramuscular fat, however, is a more objective measure of marbling content. Regression equations for each of the palatability traits utilizing intramuscular fat content accounted for a slightly greater amount of variation in consumer palatability scores than marbling content, with intramuscular fat accounting for 17%, 21%, 17%, and 17% of the variation in tenderness, juiciness, flavor, and overall palatability, respectively (Table 4).

The slope for the relationship between juiciness and intramuscular fat content was the greatest.

Table 4. Equations for linear regressions of marbling and intramuscular fat percentage and consumer palatability ratings and logistic equations for the probability of acceptable palatability ratings based on fat percentage

| Marbling score | Intercept | Regression coefficient | Adjusted $R^2$ | $P$ value |
|----------------|-----------|------------------------|----------------|-----------|
| Tenderness     | 44.82     | 0.03                   | 0.15           | <0.01     |
| Juiciness      | 44.17     | 0.03                   | 0.16           | <0.01     |
| Flavor liking  | 46.18     | 0.03                   | 0.14           | <0.01     |
| Overall liking | 43.79     | 0.03                   | 0.16           | <0.01     |
| Fat percentage |           |                        |                |           |
| Tenderness     | 51.82     | 1.60                   | 0.17           | <0.01     |
| Juiciness      | 49.35     | 1.77                   | 0.21           | <0.01     |
| Flavor liking  | 50.49     | 1.47                   | 0.17           | <0.01     |
| Overall liking | 50.28     | 1.44                   | 0.17           | <0.01     |
| Fat %—probability of acceptance | | | | |
| Tenderness     | 0.80      | 0.15                   | 0.09           | <0.01     |
| Juiciness      | 0.68      | 0.14                   | 0.08           | <0.01     |
| Flavor         | 0.86      | 0.10                   | 0.05           | <0.01     |
| Overall liking | 0.70      | 0.14                   | 0.09           | <0.01     |

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indicating that changes in fat content would result in larger increases in juiciness scores than the other traits (Figure 1). The regressions revealed that, for all traits, an increase of close to 3% intramuscular fat would be needed to increase the palatability rating by 5 units. Thompson (2004) reported a curvilinear response between fat percentage and consumer palatability scores within data from the MSA dataset. He reported changes of about 2% intramuscular fat needed for a 5-point increase (on similar 100 mm scales) in palatability ratings, with the ratings peaking at 15.5%, 17.3%, 13.8%, and 14.1% intramuscular fat for tenderness, juiciness, flavor, and overall liking, respectively. However, the author did not provide an estimate of the percentage of variation accounted for by the models ($R^2$). In this study, an evaluation of nonlinearity was assessed, but the curvilinear effect was nonsignificant ($P > 0.05$) and did not enter the model, indicating that a similar curvilinear response with palatability traits peaking at a given intramuscular fat percentage was not present in the current dataset.

Logistic regression allows for the prediction of the probability of a binomial event (yes/no) occurring using a continuous variable as an independent predictive variable. Logistic regression equations were calculated to determine the probability of a palatability trait being rated acceptable based on intramuscular fat percentage (Table 4; Figure 2). Similar to the linear regressions, models using intramuscular fat percentage to predict the probability of a sensory trait being rated acceptable were significant ($P < 0.01$), but accounted for only a small percentage (<10%) of the variation in trait acceptability. Nevertheless, the models indicated that 6.2%, 7.7%, 9.0%, and 7.6% intramuscular fat would be required for an 85% chance of tenderness, juiciness, flavor, and overall liking being rated acceptable, respectively. This corresponds to Premium choice (upper 2/3) for tenderness, juiciness, and overall liking and Prime for flavor (Savell et al., 1986). The probability increases to 90% with intramuscular fat percentages of 9.3%, 11.0%, 13.7%, and 11.0% for tenderness, juiciness, flavor, and overall liking, respectively. This indicates that at least a Prime quality grade is required for a 90% probability of each palatability trait to be rated acceptable, which is consistent with results reported in Table 3. Platter et al. (2003) also used logistic regression to predict overall beef sample acceptance based on marbling level. Those authors, similar to this study, found that marbling level represented only a small percentage (5%) of the variation in the probability of overall sample acceptance.

Using objective measures such as fat percentage to predict consumer sensory scores often produce significant relationships that account for only a small amount of variation (Dikeman, 1987, 1996). Previous authors have reported that marbling score accounted for only 27%, 20%, 26%, and 33% of the variation in trained sensory panel flavor, juiciness, tenderness, and overall palatability scores (Smith et al., 1984). Warner-Bratzler shear force values, a global industry standard, have been shown to account for only 36–73% of the variation in trained sensory panel tenderness scores (Shackelford et al., 1995; Caine et al., 2003) and 30% of the variation in consumer panel tenderness ratings (McKillip, 2017). The significant, but weak, predictive ability of both fat percentage and marbling score observed within the current dataset underscores the difficulty in using such objective measures alone to predict eating quality.

**CONCLUSIONS**

Results of this study provide a model for estimation of overall palatability based on tenderness,
juiciness, and flavor. The model allows for a comparison of the relative contribution of each of these traits to overall eating quality. Moreover, estimates related to the relative risk of overall palatability failure when one or more of the traits of tenderness, juiciness, and flavor are unacceptable were determined. The calculated risk and odds ratios provide evidence of the large impact that the failure of a single palatability trait can have on overall beef palatability. Though marbling plays a large role in beef palatability, using marbling score or fat level alone as a sole predictor of eating quality remains a challenge due primarily to the large number of factors aside from marbling impacting beef eating quality.

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