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To cite this version:
Yacine Chitour, Romain Couillet, Frédéric Pascal. On the Convergence of Maronna’s M-Estimators of Scatter. IEEE Signal Processing Letters, Institute of Electrical and Electronics Engineers, 2015, 22 (6), pp.709-712. 10.1109/LSP.2014.2367547. hal-01251753
On the convergence of Maronna’s $M$-estimators of scatter

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Abstract—In this paper, we propose an alternative proof for the uniqueness of Maronna’s $M$-estimator of scatter [1] for $N$ vector observations $y_1, \ldots, y_N \in \mathbb{R}^m$ under a mild constraint of linear independence of any subset of $m$ of these vectors. This entails in particular almost sure uniqueness for random vectors $y$, with a density as long as $N > m$. This approach allows to establish further relations that demonstrate that a properly normalized Tyler’s $M$-estimator of scatter [2] can be considered as a limit of Maronna’s $M$-estimator. More precisely, the contribution is to show that each $M$-estimator, verifying some mild conditions, converges towards a particular Tyler’s $M$-estimator. These results find important implications in recent works on the large dimensional (random matrix) regime of robust $M$-estimation.

I. INTRODUCTION

Subsequent to Huber’s introduction of robust statistics in [3], Maronna proposed in [1] a class of robust estimates for scatter matrices defined as the solution of an implicit equation. In [1], the existence and uniqueness of such a solution are proved, under conditions involving both the ratio $c_N := m/N$ of the population dimension $m$ and the sample size $N$, and the parametrization of the estimate. This constraint was largely relaxed in [4], [5]. With the recent renewed interest in robust $M$-estimation under the random matrix regime $N, m \to \infty$ with $c_N \to c_\infty \in (0, 1)$ [6]–[9], alternative proofs of existence and uniqueness have appeared motivated by this assumption of large $m$. While Maronna’s original results are valid for any (well-behaved) set of samples satisfying the condition on $c_N$, the results in e.g. [6] are expressed in probabilistic terms and are only valid for all large $m, N$.

Based on the ideas from [10]–[12], the present article proposes an alternative proof to [4] to show existence and uniqueness for all well-behaved set of samples with a known location parameter and for any $c_N \in (0, 1)$. More importantly, by a proper parametrization of the weight function appearing in Maronna’s estimator, we prove that some sequences of Maronna’s $M$-estimators converge to a unique Tyler’s distribution-free $M$-estimator of scatter [2]. This result is a novel property of the Tyler’s $M$-estimators, rigorously proved in this work. This completes the recent result (Theorem 1 of [13]) stating that the Tyler’s $M$-estimator is the Maximum Likelihood estimator (MLE) of the scatter for various complex elliptically symmetric (CES) distributions as well as for the angular central Gaussian (ACG) distributions [14].

The paper is organized as follows: Section II presents our main results as well as Monte-Carlo simulations that corroborate our theoretical claims, the proofs of which are provided in Section III. Section IV draws some conclusions and perspectives of this work.

II. NOTATIONS AND STATEMENT OF THE RESULTS

Let $R_+$ (resp. $R_+^*$) be the (resp. strictly) positive real line. We use $M_m(R)$ and $\text{Sym}_m$ to denote the vector space of $m \times m$ matrices with real entries and the linear subspace of $M_m(R)$ made of the symmetric matrices, respectively. We also use $\text{Sym}_m^+$ and $\text{PSD}_m$ to denote the non trivial cones in $M_m(R)$ of the non negative symmetric matrices and of the symmetric positive definite matrices, respectively. Also, $(\cdot)^T$ stands for the transpose, $\text{Tr}(\cdot)$ and $\text{det}(\cdot)$ for the trace and the determinant. On $M_m(R)$, we use the inner product defined by the Frobenius norm $\|A\| = \sqrt{\text{Tr}(AA^T)}$. We also use $\leq$ to denote the partial order on $\text{Sym}_m$ and $I_m$ the $m \times m$ identity matrix. Functions of two non negative real variables $(t, x)$ will be considered. If $f$ is such a function, we use $f_t$, $f_x$, $f_{tx}$, . . . to denote (when defined) the partial derivatives of $f$ with respect to $t$ and/or $x$.

Definition II.1 A family $(y_i)_{1 \leq i \leq N}$ of vectors in $\mathbb{R}^m$ is admissible if

(C1) for $1 \leq i \leq N$, $\|y_i\| = 1$;
(C2) the vectors in any subset of size $m$ of \{y_1, \cdots, y_N\} are linearly independent

This definition straightforwardly implies that if $(y_i)_{1 \leq i \leq N}$ is an admissible family of vectors in $\mathbb{R}^m$ and if $m$ vectors (say) $y_1, \cdots, y_m$ which are then linearly independent by (C2) are fixed, for $m + 1 \leq l \leq N$, we can write $y_l = \sum_{j=1}^{m} \gamma_j y_j$. Then, $\gamma_j \neq 0$ for every $1 \leq j \leq m$ and $m + 1 \leq l \leq N$. 

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Let us now consider maps $u : (\mathbb{R}^+)^2 \to \mathbb{R}^+$ of class $C^1$ satisfying:

(U1) $u(t, \cdot)$ is strictly decreasing;
(U2) for every $t > 0$, $v(t, x) := x \mapsto xu(t, x)$ is increasing on $\mathbb{R}^+$ and $l_t := \sup_{x \geq 0} v(t, x) > m$.

We furthermore define, for every $x > 0$, $u(0, x) = \frac{m}{x}$. Note that, by continuity of $u$, $\forall x > 0$, $\lim_{t \to 0^+} u(t, x) = m$. Also, according to (U1) and (U2), for each $t, x > 0$, 

$$v(t, x) = m + tv_1(x) + tw(t, x),$$

with $v_1(\cdot) := v_t(0, \cdot)$ and $\forall x > 0$, $\lim_{t \to 0^+} w(t, x) = 0$. By a simple computation, one has that $v_1$ is a nondecreasing function on $\mathbb{R}^+$. For further use, we introduce the following additional notation. Let $x_t > 0$ be the unique positive number such that, $\forall t > 0$, $v(t, x_t) = x_t u(t, x_t) = m$.

We further consider the following assumption

(U3) \[
\begin{cases} 
  v_x := dv/dx > 0 \\
  v_1 \text{ is increasing} \\
  0 < \liminf_{t \to 0^+} x_t \leq \limsup_{t \to 0^+} x_t < \infty.
\end{cases}
\]

If the latter occurs and $u$ is of class $C^2$, then $w(t, x) = tw_1(x) + o(t)$, with $w_1(\cdot) := w_t(0, \cdot)$ continuous on $(\mathbb{R}^+)^2$, the convergence in (U2) is uniform in $x$ on any compact of $\mathbb{R}^+$ and $x_t$ converges to the unique solution $x_0$ of $v_1(x) = 0$.

We use $\tilde{u}(t, x)$ to denote the particular function

$$\tilde{u}(t, x) = \frac{m(1 + t)}{x + t}$$

which is analytic on every compact of $(\mathbb{R}^+)^2 \setminus \{(0, 0)\}$. Moreover, $\tilde{l}_t = m(1 + t)$, $\tilde{v}_1(x) = m(1 - \frac{1}{x})$ and $\tilde{w}(t, x) = \frac{-mt}{x + t}$.

The objective of the work is to study the solutions of the equation given, for all $t > 0$, by

$$(E\!q)_t \quad M = \frac{1}{N} \sum_{i=1}^{N} u(t, y_i^T M^{-1} y_i) y_i y_i^T.
$$

and to characterize them in the limit where $t \to 0$. Taking into account our definitions, if a solution to $(E\!q)_t$ exists, it must belong to $\text{PSD}_m$.

Remark II.3

1) The interest of Theorem II.2, in addition to providing an alternative proof for the existence and uniqueness, lies in the convergence of all $M$-estimators to a Tyler’s $M$-estimator. This limit can be different (by a scale factor) from one $M$-estimator to another. While this result was expected, this paper rigorously proves it.

2) Moreover, the theorem provides a way of understanding why the Tyler’s estimator is the outmost robust $1$ $M$-estimator. Indeed, considering a ML approach, the weight function $u(t, x)$ is derived from the observations probability density function (PDF) and in such a case, $t \to 0$ means that the underlying distribution becomes more and more heavy-tailed. For instance, considering $t$ as the exponent parameter of a Generalized Gaussian distribution or of a W-distribution, the smaller the value of $t > 0$ is, the heavier-tailed is the distribution. This is also the case for the degree of freedom of a Student-t distribution or the shape parameters of a K-distribution or of a Compound-Gaussian with inverse Gaussian texture (see [14] for more details). In all these cases, the MLEs satisfy the assumptions of Theorem II.2 (at least for small values of $t$) and should be more robust when the distribution is heavier-tailed.

To summarize, this result theoretically motivates the use of the Tyler’s estimator, since it will perform similarly as MLEs in heavy-tailed distribution contexts.

\[1\] Here the robustness has to be understood as the classical property considered in the robust estimation theory literature, see e.g. [15].
To illustrate Theorem II.2, Figure 1 presents the mean square error \( C(t) \equiv E[|M(t) - M_0|^2] \) between Tyler’s \( M \)-estimator and the Student-t MLE versus the parameter \( t \), called the degree of freedom of the multivariate Student-t distribution [14], defined through the weight function \( u(t, x) = \frac{m+x}{m+x+t} \). We take here \( N = m + 1 = 51 \). The data are zero-mean Gaussian distributed with Toeplitz covariance matrix, the \((i, j)\) entry of which is equal to \( \rho^{i-j} \), for some \( \rho \in (0, 1) \). As proved in Theorem II.2, Item (A) is illustrated in the case where \( N = m + 1 \) while Item (B) is illustrated for the Student-t MLE for different population covariance matrices.

III. PROOF OF THEOREM II.2

The strategy of the proof is as follows: for every \( t > 0 \), we first build a positive functional \( H(t, \cdot) \) over \( \text{PSD}_m \) whose critical points (if any) are exactly the solutions of (Eq.)\(_t\). To establish the existence of such critical points, we show that \( H(t, \cdot) \) is uniformly bounded and tends to zero at the boundary of \( \text{PSD}_m \). To obtain uniqueness, we show that solutions of (Eq.)\(_t\) are all local strict maxima of \( H(t, \cdot) \) and conclude by applying the mountain pass theorem (cf. [16]). This gives Item (A). Item (B) is then obtained using the implicit function theorem and some limiting arguments.

For \( t > 0 \), we define the function

\[
    h : \mathbb{R}_+^2 \times \mathbb{R}_+ \rightarrow \mathbb{R}_+^*
    (t, x) \rightarrow e^{-\frac{1}{3} \int_x^y u(t, y) \, dy}.
\]

Then \( -\frac{h}{t^2} = \frac{h}{t} \) and \( h(t, x_i) = 1 \). Set \( \bar{h}(0, x) = \frac{1}{x} \) for \( x > 0 \) and \( g : \mathbb{R}_+^* \times \mathbb{R}_+ \rightarrow \mathbb{R}_+^* \) with \( g(t, x) = xh(t, x) \).

In the case where \( u = \bar{u}, \forall(t, x) \in (\mathbb{R}_+)^2 \setminus \{(0, 0)\}, x_i \equiv 1, \bar{h}(t, x) = \left(1 + t + \frac{1}{x + t}\right)^m, \bar{g}(t, x) = x \left(1 + t + \frac{1}{x + t}\right)^m.\]

Then, define the functional \( H(t, \cdot) \) as

\[
    H : \mathbb{R}_+^2 \times \text{PSD}_m \rightarrow \mathbb{R}_+^*
    (t, M) \rightarrow \frac{\prod_{i=1}^N h(t, y_i^T M^{-1} y_i)^m}{(\det M)^N}.
\]

as well as the functional considered in [10]

\[
    B : \text{PSD}_m \rightarrow \mathbb{R}_+^*
    M \rightarrow \frac{\prod_{i=1}^N h(0, y_i^T M^{-1} y_i)^m}{(\det M)^N}.
\]

Lemma III.1 For \( t > 0 \) and \( M \in \text{PSD}_m \), one has \( -\Delta H_x(t, M)M/NH(t, M) = M - \frac{1}{N} \sum_{i=1}^N u(t, y_i^T M^{-1} y_i) \), with \( H_x(t, M) \) the gradient of \( H(t, \cdot) \). In particular, \( M \) is a solution of (Eq.)\(_t\) if and only if \( M \) is a critical point of \( H(t, \cdot) \).

Lemma III.2 \( \forall t > 0, M \in \text{PSD}_m, H(t, M) \leq B(M). \) As a consequence, \( \lim_{M \rightarrow \partial \text{PSD}_m} H(t, M) = 0 \), so that \( H(t, \cdot) \) admits critical points.

Proof of Lemma III.2. An immediate calculus yields that \( x \mapsto g(t, x) \) reaches its maximum 1 at \( x = x_t \). As a consequence, for \( t > 0, M \in \text{PSD}_m, H(t, M) \leq B(M) \). Moreover, \( \lim_{x \rightarrow 0, \infty} g(t, x) = \lim_{x \rightarrow 0, \infty} xh(t, x) = 0 \). For the limit at \( x = 0, \) this is obvious. For \( x \rightarrow \infty, \) note that \( \ln(g(t, x)) = \frac{1}{m} \int_{x}^{+\infty} \frac{m - y u(t, y)}{y} \, dy \) and, since \( m - l_i < 0 \), it is equivalent to \((m - l_i)\lim(x) \) as \( x \rightarrow \infty \). Consider now a sequence \( (M_k)_{k \geq 0} \) in \( \text{PSD}_m \) converging to \( \partial \text{PSD}_m \). For \( k \geq 0 \), set \( M_k = \rho_k N_k \) with \( \rho_k = \|M_k\| \) and \( N_k = \frac{M_k}{\rho_k} \). Note that \( \partial \text{PSD}_m \) is made of matrices either non invertible or with norm going to infinity. Therefore, up to subsequences, either (i) \( (N_k)_{k \geq 0} \) converges itself to \( \partial \text{PSD}_m \) or (b) the sequence \( (\rho_k)_{k \geq 0} \) converges to zero or infinity and there exists \( \exists \alpha > 0, \forall k \geq 0, N_k \geq \alpha I_m \). If Case (i) occurs, then \( \forall k \geq 0, H(t, M_k) \leq B(N_k) \), which tends to zero as \( k \rightarrow \infty \) (cf. [10]). In Case (ii),

\[
    H(t, M_k) = \frac{\prod_{i=1}^N h(t, x_{i,k})^m}{\rho_k^N \det(N_k)^N} = B(N_k) \prod_{i=1}^N g(t, x_{i,k})^m
\]

where \( x_{i,k} = y_i^T N_k^{-1/2} y_i/\rho_k \). As \( k \rightarrow \infty, x_{i,k} \) tends either to zero or infinity and we conclude. For \( t > 0, H(t, \cdot) \) is uniformly bounded over \( \text{PSD}_m \) since \( B(\cdot) \) is. So \( H(t, \cdot) \) has a global maximum which must belong to \( \text{PSD}_m \) since \( H(t, M) \rightarrow 0 \) as \( M \) tends to the boundary of \( \text{PSD}_m \). So \( H(t, \cdot) \) admits critical points.

Lemma III.3 Let \( t > 0 \). Then all critical points of \( H(t, \cdot) \) are local strict maxima.

![Fig. 1. Convergence of \( M(t) \) towards \( M_0 \) when \( t \rightarrow 0 \) for \( N = m + 1 = 51 \). The criterion used is the MSE: \( C(t) = E[|M(t) - M_0|^2] \).](image)
Proof of Lemma III.3. We show that, if $M$ is a critical point then the Hessian of $H(t, \cdot)$ at $M$ is a negative definite quadratic form implying that $M$ is a local strict maximum of $H(t, \cdot)$. Let $M \in \text{PSD}_m$ be a critical point of $H(t, \cdot)$. Then, one gets that for every $Q \in \text{Sym}_m$,

\[
\langle Q, \text{Hess}_M(Q) \rangle = -NH(t, M) \left[ \langle Q, M^{-1}Q M^{-1} \rangle \right.
+ \frac{1}{N} \sum_{i=1}^{N} u_x(t, y_i^T M^{-1} y_i)(y_i^T M^{-1} y_i)^2 \left. \right]
\]

Let $R := M^{-1/2}QM^{-1/2}$ and $d_i := M^{-1/2}y_i$, one has

\[
- \frac{\langle Q, \text{Hess}_M(Q) \rangle}{NH(t, M)} = \|R\|^2 + \frac{1}{N} \sum_{i=1}^{N} u_x(t, \|d_i\|^2)(d_i^T R d_i)^2.
\]

Recall that $M$ is a critical point of $H(t, \cdot)$ and thus a solution of (Eq)$_t$, i.e.,

\[
I_m = \frac{1}{N} \sum_{i=1}^{N} u(t, \|d_i\|^2)d_i d_i^T.
\]

Multiplying (8) by $R$ on both left and right, taking the trace and plugging the result into (7) gives

\[
(7) = \frac{1}{N} \sum_{i=1}^{N} u(t, \|d_i\|^2)\|R d_i\|^2 + u_x(t, \|d_i\|^2)(d_i^T R d_i)^2. \tag{7}
\]

Let $I_Q = \{i \in \{1, \cdots, N\}, Rd_i \neq 0\}$. Then

\[
(7) = \frac{1}{N} \sum_{i \in I_Q} \|R d_i\|^2 \left[ u(t, \|d_i\|^2) + \|d_i\|^2 u_x(t, \|d_i\|^2) r_i \right]
\]

where $r_i := (d_i^T R d_i/\|d_i\|\|R d_i\|)^2$. Using $0 \leq r_i \leq 1$ (by Cauchy-Schwarz’s inequality) and $u_x \leq 0$ (since $u$ is of class $C^1$ and verifies (U1)), we have $r_i u_x(\cdot, \cdot) \geq u_x(\cdot, \cdot)$. Then, recalling that $v(t, x) = xu(t, x)$,

\[
(7) \geq \frac{1}{N} \sum_{i \in I_Q} \|R d_i\|^2 v_x(t, \|d_i\|^2) \geq 0.
\]

Moreover, if $Q \neq 0, I_Q \neq \emptyset$ and there exists $\bar{i}$ such that $v_x(t, \|d_{\bar{i}}\|^2) > 0$. Therefore $\langle Q, \text{Hess}_M(Q) \rangle < 0$, i.e., $\text{Hess}_M$ is negative definite, concluding the proof. \qed

Lemma III.4 Let $t > 0$. Then (Eq)$_t$ admits a unique solution, $M(t)$, the unique strict maximum of $H(t, \cdot)$.

Proof of Lemma III.4. We reason by contradiction assuming $H(t, \cdot)$ admits at least two local strict maxima. Applying the mountain-pass theorem [16] to the functional $1/H(t, \cdot)$ which tends to infinity in the vicinity of $\partial \text{PSD}_m$, we obtain the existence of a saddle point of $F$ in $\text{PSD}_m$ which is contradictory to Lemma III.3.

We next prove that $M(t)$ is uniformly bounded in $\text{PSD}_m$ as $t \to 0$, i.e.

**Lemma III.5 There exists $0 < a \leq b$ and $t_0 > 0$ such that, for every $t \in (0, t_0)$, $a I_m \leq M(t) \leq b I_m$.**

Proof of Lemma III.5. Let $P$ be the unique matrix of $\text{PSD}_m$ satisfying $B(P) = \max_{M \in \text{PSD}_m} B(M)$ and $\text{Tr}(M) = m$. Then, for every $t > 0$, $H(t, P) \leq H(t, M(t))$ and $B(M(t)) \leq B(P)$. Multiplying both inequalities, after simplifications, we get $\prod_{i=1}^{N} g(t, y_i^T P^{-1} y_i) \leq \prod_{i=1}^{N} g(t, y_i^T M(t)^{-1} y_i) \leq 1$, with $\prod_{i=1}^{N} g(t, y_i^T P^{-1} y_i) \to 1$ as $t \to 0$. So there exists $t_0 > 0$ such that, for every $t \in (0, t_0)$ and $1 \leq i \leq N, 1/2 \leq g(t, y_i^T M(t)^{-1} y_i)$, and, since (U3) holds true, there exists $0 < a \leq b$ s.t. for every $t \in (0, t_0)$ and $1 \leq i \leq N, u(t, y_i) \leq u(t, y_i M(t)^{-1} y_i) \leq u(t, a)$, hence $u(t, b)C \leq M(t) \leq u(t, a)C$ with $C := \frac{a}{b} \sum_{i=1}^{N} y_i y_i^T$. One concludes easily.

**Lemma III.6 Under the conditions of Theorem II.2, \lim_{t \to 0} M(t) = M_0 solution of (Eq)$_0$ given by $M_0 = \xi_u P$, where $\xi_u > 0$ is the unique solution of (3).**

Proof of Lemma III.6. Since $M(\cdot)$ is uniformly bounded in $\text{PSD}_m$ as $t \to 0$, its accumulation points still belong to $\text{PSD}_m$ and are necessarily of the form $\mu P$ where $\mu > 0$ and $P$ is the solution of (Eq)$_0$ with trace $m$. Taking the trace in (8), one gets $m = \frac{1}{N} \sum_{i=1}^{N} v(t, \|d_i(t)\|^2)$, where $d_i(t) = M(t)^{-1/2} y_i$ for $1 \leq i \leq N$. Using (1) and (U3), one deduces that, for every $t > 0, \sum_{i=1}^{N} v_1(\|d_i(t)\|^2) + t \sum_{i=1}^{N} v_1(\|d_i(t)\|^2) + o(t) = 0$. Consider an accumulation point $\mu P$ of $M(\cdot)$ as $t \to 0$. Then, up to a subsequence, $\lim_{t \to 0} M(t) = \mu P$ and, for $1 \leq i \leq N$, $\lim_{t \to 0} d_i(t) = P^{-1/2} y_i/\sqrt{\mu}$. According to (U3), the second sum in the previous equation tends to zero as $t \to 0$ and we are left with $\sum_{i=1}^{N} v_1(\|d_i(t)\|^2)/\sqrt{\mu} = 0$. Since the left-hand side of the latter defines a decreasing function of $\mu$, it has a unique solution denoted $\xi_u > 0$, which concludes the proof since $M(\cdot)$ admits a unique accumulation point as $t \to 0$. \qed

**IV. CONCLUSIONS**

In this paper, an alternative proof for existence and uniqueness for the Maronna’s $M$-estimators is provided. More importantly, using this particular approach leads to draw some connections between Maronna’s and Tyler’s estimators by expressing (properly scaled) Tyler’s estimator in terms of a limit of a class of Maronna’s estimators. This result may also find interest in studies of Tyler’s $M$-estimator in the large random matrix regime.
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