A rank-based selection with cardinal payoffs and a cost of choice

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Abstract

A version of the secretary problem is considered. The ranks of items, whose values are independent, identically distributed random variables $X_1, X_2, \ldots, X_n$ from a uniform distribution on $[0; 1]$, are observed sequentially by the grader. He has to select exactly one item, when it appears, and receives a payoff which is a function of the unobserved realization of random variable assigned to the item diminished by some cost. The methods of analysis are based on the existence of an embedded Markov chain and use the technique of backward induction. The result is a generalization of the selection model considered by Bearden (2006). The asymptotic behaviour of the solution is also investigated.

Key words: optimal stopping, sequential search, secretary problem, rank-based selection, cardinal payoffs, Markov chain,

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1 Introduction

Although a version of the secretary problem (the beauty contest problem, the dowry problem or the marriage problem) was first solved by Cayley (1875), it was not until

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five decades ago there had been sudden resurgence of interest in this problem. Since the articles by [Gardner 1960a,b] the secretary problem has been extended and generalized in many different directions by [Gilbert and Mosteller 1966]. Excellent reviews of the development of this colourful problem and its extensions have been given by [Rose 1982b], [Freeman 1983], [Samuels 1991] and [Ferguson 1989]. The classical secretary problem in its simplest form can be formulated following [Ferguson 1989]. He defined the secretary problem in its standard form to have the following features:

(i) There is only one secretarial position available.
(ii) The number of applicants, \( N \), is known in advance.
(iii) The applicants are interviewed sequentially in a random order.
(iv) All the applicants can be ranked from the best to the worst without any ties. Further, the decision to accept or to reject an applicant must be based solely on the relative ranks of the interviewed applicants.
(v) An applicant once rejected cannot be recalled later. The employer is satisfied with nothing but the very best.
(vi) The payoff is 1 if the best of the \( N \) applicants is chosen and 0 otherwise.

This model can be used as a model of choice in many decisions in everyday life, such as buying a car, hiring an employee, or finding an apartment (see [Corbin 1980]). The part of research has been devoted to modified version of the problem where some important assumption of the model has been changed to fit it to the real life context. There are analysis of decision maker’s aims. It could be that he will be satisfied by choosing one of the \( K \) best (see [Gusein-Zade 1966], [Frank and Samuels 1980]). It was shown that the optimal strategy in this problem has very simple threshold form. The items are observed and rejected up to some moments \( j_r \) (thresholds) after which it is optimal to accept the first candidate with relative rank \( r \), \( r = 1, 2, \ldots, K \). The thresholds \( j_r \) are decreasing on \( r \). This strategy is rather intuitive. When the candidates run low we admit acceptance the lowest rank of chosen item. If the aim is to choose the second best item then the form of the optimal strategy is not so intuitively obvious (see [Szajowski 1982], [Rose 1982a], [Mori 1985]). In the same time the possibility of backward solicitation and uncertain employment was also investigated (see [Yang 1974], [Smith and Deely 1975], [Smith 1975]).

There are also experimental research with subjects confronted with the classical secretary problem (see [Seale and Rapoport 1997, 2000]). The optimal strategy of the grader in the classical secretary problem is to pass \( k^*_N - 1 \) applicants, where \( k^*_N \approx [Ne^{-1}] \) and stop at the first \( j \geq k^*_N \) which is better that those seen so far. If none exists nothing is chosen. The experimental study by [Seale and Rapoport 1997] of this problem shows that subjects under study have tendency to terminate their search earlier than in the optimal strategy. [Bearden 2006] has considered application the best choice problem to the model of choice for the trader who makes her selling decision at each point in time solely on the basis of the rank of the current price with respect to the previous prices, but, ultimately, derive utility from the true value of the selected observation and not from its rank. The
assumption (vi) is not fulfilled in this case. Bearden (2006) has made efforts to explain this effect and the new payoff scheme has proposed. He shows that if the true values $X_j$ are i.i.d. uniformly distributed on $[0, 1]$ then for every $N$ the optimal strategy is to pass $c - 1$ applicants, and stop with the first $j \geq c$ with rank 1. If none exists, stop at time $N$. The optimal value of $c$ is either $\lfloor \sqrt{N} \rfloor$ or $\lceil \sqrt{N} \rceil$.

This payoff scheme when the i.i.d. $X_j$’s come from other than the uniform distribution has been studied by Samuel-Cahn (2005). Three different families of distributions, belonging to the three different domains of attraction for the maximum, have been considered and the dependence of the optimal strategy and the optimal expected payoff has been investigated. The different distributions can model various tendency in perception of the searched items.

In this paper the idea of payoff function dependent on the true value of the item is modified to include the different personal costs of choice of the item. The cost of observation in the secretary problem with payoffs dependent on the real ranks has been investigated by Bartoszynski and Govindarajulu (1978) (see also Yeo (1998)). However, the cost of decision is different problem than the cost of observation. It will be shown that the optimal number of items one should skip is a function of this personal cost. At the last moment the payoff function can be slightly differently defined than in Bearden (2006)’s paper. The asymptotic expected return and asymptotic behaviour of the optimal strategy will be studied.

The organization of the paper are as follows. In Section 2 the related to the secretary problem Markov chain is formulated. This section is based mainly on the suggestion from Dynkin and Yushkevich (1969) and the results by Szajowski (1982) and Suchwalko and Szajowski (2002). In the next sections the solution of the rank-based secretary problem with cardinal payoff and the personal cost of grader is given. In Section 3 the exact and asymptotic solution is provided for the model formulated in Section 2. In this consideration the asymptotic behaviour of the threshold defining the optimal strategy of the grader is studied.

In the last section the comparison of obtained results are given.

2 Mathematical formulation of the model

Let us assume that the grader observes a sequence of up to $N$ applicants whose values are i.i.d. random variables $\{X_1, X_2, \ldots, X_N\}$ with uniform distribution on $\mathbb{E} = [0, 1]$. The values of the applicants are not observed. Let us define

$$R_k = \#\{1 \leq i \leq k : X_i \leq X_k\}.$$
The random variable $R_k$ is called relative rank of $k$-th candidate with respect to the moment $k$. The grader can see the relative ranks instead of the true values. All random variables are defined on a fixed probability space $(\Omega, \mathcal{F}, \mathbb{P})$. The observations of random variables $R_k$, $k = 1, 2, \ldots, N$, generate the sequence of $\sigma$-fields $\mathcal{F}_k = \sigma\{R_1, R_2, \ldots, R_k\}$, $k \in \mathbb{T} = \{1, 2, \ldots, N\}$. The random variables $R_k$ are independent and $\mathbb{P}\{R_k = i\} = \frac{1}{k}$.

Denote by $\mathcal{M}^N$ the set of all Markov moments $\tau$ with respect to $\sigma$-fields $\{\mathcal{F}_k\}_{k=1}^N$. Let $q : \mathbb{T} \times \mathbb{S} \times \mathbb{E} \to \mathbb{R}^+$ be the gain function. Define

(1) \[ v_N = \sup_{\tau \in \mathcal{M}^N} \mathbb{E}q(\tau, R_\tau, X_\tau). \]

We are looking for $\tau^* \in \mathcal{M}^N$ such that $\mathbb{E}q(\tau^*, R_{\tau^*}, X_{\tau^*}) = v_N$.

Since $\{q(n, R_n, X_n)\}_{n=1}^N$ is not adapted to the filtration $\{\mathcal{F}_n\}_{n=1}^N$, the gain function can be substituted by the conditional expectation of the sequence with respect to the filtration given. By property of the conditional expectation we have

\[
\mathbb{E}q(\tau, R_\tau, X_\tau) = \sum_{r=1}^N \int_{\{\tau = r\}} q(\tau, R_\tau, X_\tau) d\mathbb{P}
= \sum_{r=1}^N \int_{\{\tau = r\}} \mathbb{E}[q(r, R_r, X_r)|\mathcal{F}_r] d\mathbb{P}
= \mathbb{E}\tilde{g}(\tau, R_\tau),
\]

where

(2) \[ \tilde{g}(r, R_r) = \mathbb{E}[q(r, R_r, X_r)|\mathcal{F}_r] \]

for $r = 1, 2, \ldots, N$. On the event $\{\omega : R_r = s\}$ we have $\tilde{g}(r, s) = \mathbb{E}[q(r, R_r, X_r)|R_r = s]$.

**Assumption 1** In the sequel it is assumed that the grader wants to accept the best so far applicant.

The function $\tilde{g}(r, s)$ defined in (2) is equal to 0 for $s > 1$ and non-negative for $s = 1$. It means that we can choose the required item at moments $r$ only if $R_r = 1$. Denote $h(r) = \tilde{g}(r, 1)$.

The risk is connected with each decision of the grader. The personal feelings of the risk are different. When the decision process is dynamic we can assume that the feeling of risk appears randomly at some moment $\xi$. Its distribution is a model of concern for correct choice of applicant.

**Assumption 2** It is assumed that $\xi$ has uniform distribution on $\{0, 1, \ldots, N\}$. 

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Remark 2.1 Let us assume that the cost of choice or the measure of stress related to the decision of acceptance of the applicant is $c$. It appears when the decision is after $\xi$ and its measure will be random process $C(t) = cI_{\{t \geq \xi\}}$. Based on the observed process of relative ranks and assuming that there are no acceptance before $k$ we have

$$c(k, t) = \mathbb{E}[C(t) | F_k] = c \frac{N - t + 1}{N - k + 1}.$$  

The applied model is a consequence of observation that the fear of the wrong decision today is highest than the concern for the consequence of the future decision.

Assumption 3 The aim of the grader is to maximize the expected value of applicant chosen and at the same time to minimize the cost of choice.

In this case the function

$$q(t, R_t, X_t) = g_c(t, R_t, X_t) = \begin{cases} (X_t - C(t))I_{\{R_t = 1\}}(R_t) & \text{if } t < N, \\ X_N - c & \text{otherwise.} \end{cases}$$

Since $X_t$ are i.i.d. random variables with the uniform distribution on $[0, 1]$ we have for $t \geq r$

$$\tilde{g}_c(r, t, R_t) = \mathbb{E}[g_c(t, R_t, X_t) | F_r] = (\frac{t}{t + 1} - c \frac{N - t + 1}{N - r + 1})I_{\{R_t = 1\}}(R_t)$$

(see Resnick (1987)). Let us denote $\tilde{h}(r, s) = \tilde{g}(r, s, 1)$.

Define $W_0 = 1$, $\gamma_t = \inf \{r > \gamma_{t-1} : Y_r = 1\}$ (inf $\emptyset = \infty$) and $W_t = \gamma_t$. If $\gamma_t = \infty$, then define $W_t = \infty$. $W_t$ is the Markov chain with following one step transition probabilities

$$p(r, s) = \mathbb{P}\{W_{t+1} = (s, 1) | W_t = (r, 1)\} = \begin{cases} \frac{1}{s}, & \text{if } r = 1, s = 2, \\ \frac{r}{s(s-1)}, & \text{if } 1 < r < s, \\ 0, & \text{if } r \geq s \text{ or } r = 1, s \neq 2, \end{cases}$$

with $p(\infty, \infty) = 1$, $p(r, \infty) = 1 - \sum_{s=r+1}^{N} p(r, s)$. Let $\mathcal{G}_t = \sigma\{W_1, W_2, \ldots, W_t\}$ and $\tilde{\mathcal{M}}^N$ be the set of stopping times with respect to $\{\mathcal{G}_t\}_{t=1}^{N}$. Since $\gamma_t$ is increasing, then we can define $\tilde{\mathcal{M}}^N_{t+1} = \{\sigma \in \tilde{\mathcal{M}}^N : \gamma_\sigma > t\}$.

Let $\mathbb{P}_\sigma(\cdot)$ be probability measure related to the Markov chain $W_t$, with trajectory starting in state $r$ and $\mathbb{E}_\sigma(\cdot)$ the expected value with respect to $\mathbb{P}_\sigma(\cdot, 1)$. From (6) we can see that the transition probabilities depend on moments $r$ where items with relative rank 1 appears.
Taking into account the form of the payoff function \(g\), the two-dimensional Markov chain should be considered. Denote \(Z_t : \Omega \rightarrow \mathbb{T} \times \mathbb{T}\) the Markov chain with the following one step transition probabilities

\[
P(Z_{t+1} = (s, j) | Z_t = (s, i)) = \begin{cases} \frac{i}{j(j - 1)} & \text{for } s < i < j \leq N, \\ \frac{s}{k(k - 1)} & \text{for } s < k < i \leq N, \\ 0 & \text{otherwise.} \end{cases}
\]

and 0 otherwise.

Let us introduce the operators based on (7) and (6)

\[
T_\tilde{h}(r, s) = \mathbb{E}_{(r, s)} \tilde{h}(Z_1) = \sum_{j=s+1}^{N-1} \frac{s}{j(j - 1)} \tilde{h}(r, j) + \left(1 - \sum_{j=s+1}^{N-1} \frac{s}{j(j - 1)}\right) \left(\frac{1}{2} - c\right),
\]

\[
T_h(r) = \mathbb{E}_r h(W_1) = \sum_{j=r+1}^{N-1} \frac{r}{j(j - 1)} \tilde{h}(r, j) + \left(1 - \sum_{j=r+1}^{N-1} \frac{r}{j(j - 1)}\right) \left(\frac{1}{2} - c\right).
\]

3 The cost of fear in the rank-based secretary problem with cardinal value of the item.

Let \(\mathcal{M}_r^N = \\{\tau \in \mathcal{M}^N : r \leq \tau \leq N\}\) and \(v_N(r) = \sup_{\tau \in \mathcal{M}_r^N} \mathbb{E} g_c(\tau, R_\tau, X_\tau)\). The following algorithm allows to construct the value of the problem \(v_N\). Let

\[
v_N(N) = \mathbb{E} g_c(N, R_N, X_N) = \mathbb{E}(X_N) - c.
\]

and for \(r < N\)

\[
w_N(r, s) = \max\{h(r, s), Tw_N(r, s)\},
\]

\[
v_N(r) = \max\{h(r), Tv_N(r)\}.
\]

One can consider the stopping sets

\[
\Gamma_r = \{(r, s) : h(r, s) \geq w_N(r, s), \ r < s\} \cup \{(r, N)\},
\]

\(r \in \mathbb{T}\). In class of such stopping sets there are solutions of restricted problem. Based on this partial solution the optimal stopping time is constructed and it is shown that \(v_N = v_N(1)\).

Lemma 3.1 For the considered problem with the payoff function \(g\) and \(c \in \mathbb{R}^+\), there is \(k_0\) such that for \(r \geq k_0\) the optimal stopping time \(\tau^*\) in \(\mathcal{M}_r^N\) has a form \(\tau^* = \inf\{s \geq r: Y_s = 1\}\) i.e. the stopping set is \(\Gamma_r = \{(r, s) : s \geq r, Y_r = 1\} \cup \{(r, N)\}\).
PROOF. The function $\tilde{h}(k, r) = \frac{r}{r+1} - c \frac{N-r+1}{N-k+1}$ is increasing on $r \geq k$. For $r = N$ we have $w_N(k, N) = \frac{1}{2} - c$. Let us construct the one step look ahead stopping time and let us define $k_0 = \min\{1 \leq k \leq N : h(s) \geq Th(s) \text{ for every } s \in [k, N]\}$. For $j \geq k \geq k_0$ we have $h(k) \leq h(j) \leq \tilde{h}(k, j)$ and by definition of $k_0$ we have $\tilde{h}(k, j) \geq h(k) \geq Th(k) \geq Th(k, j)$. The value of the problem $w_N(k, r) = \tilde{h}(k, r)$ and the optimal stopping time on $\mathcal{M}_k^N$ is defined by the stopping set $\Gamma \cap k_0$. Therefore we have $Tv_N(r) = Th(r)$ for $r \geq k_0$ and the one step look ahead rule is optimal in $\mathcal{M}_k^N$.

$\blacksquare$

Remark 3.2 Let us assume that $s > k > k_0$. We take limits of $k \rightarrow y$ and $s \rightarrow x$ as $N \rightarrow \infty$. We get

$$\tilde{h}(y, x) = \lim_{N \rightarrow \infty, \frac{k}{N} \rightarrow y; \frac{s}{N} \rightarrow x} \tilde{h}(k, s) = 1 - c \frac{1-x}{1-y}$$

$$\tilde{h}(y, x) = \lim_{N \rightarrow \infty, \frac{k}{N} \rightarrow y; \frac{s}{N} \rightarrow x} Th(k, s) = 1 - \frac{x}{2} - cx - \frac{1-x}{1-y} c - \frac{xc}{1-y} \log(x).$$

For $c \in (0, +\infty)$ the equation $\log(y) = (y - 1)(\frac{1}{2x} + 1)$ has one root $\alpha \in (0, 1)$. When $x \geq y \geq \alpha$ then $\tilde{h}(y, x) \leq \tilde{h}(y, x)$.

The optimal stopping time $\tau^*$ is defined as follows: one have to stop at the first moment $r$ when $y_r = 1$, unless $v_N(r) > h(r)$. We can define the stopping set $\Gamma = \{r : h(r) \geq v_N(r)\} \cup \{N\}$.

Theorem 3.3 For every $c \in [0, +\infty)$ there is $k_0$ such that $\Gamma = \{r : r \geq k_0, Y_r = 1\} \cup \{N\}$ and $v_N = v_N(k_0 - 1)$.

PROOF. The function $h(r) = \frac{r}{r+1} - c$ is increasing on $r$. For $r = N$ we have $v_N(N) = \frac{1}{2} - c$. Let us construct the one step look ahead stopping time and let us define $k_0 = \min\{1 \leq k \leq N : h(s) \geq Th(s) \text{ for every } s \in [k, N]\}$. For $j \geq k \geq k_0$ we have $h(k) \leq h(j) \leq \tilde{h}(k, j)$ and by definition of $k_0$ the value of the problem on $\mathcal{M}_k^N$ is equal to $v_N(k_0 - 1) = Th(k_0 - 1)$ and the one step look ahead rule is optimal in this set of stopping times. For $r \leq k_0 - 1$ we have $h(r) \leq v_N(k_0 - 1)$. If we do not stop at the moment $r < k_0 - 1$ we get
\[ v_N(r) = \sum_{j=r+1}^{k_0-1} \frac{k_0-1}{j(j-1)} v_N(k_0-1) \]
\[ + \frac{r}{k_0-1} \left( \sum_{j=k_0}^{N-1} \frac{k_0-1}{j(j-1)} \tilde{h}(k_0-1, j) + \left( 1 - \sum_{j=k_0}^{N-1} \frac{k_0-1}{j(j-1)} \right) \frac{1}{2} \right) \]
\[ = rv_N(k_0-1)\left( \frac{1}{r} - \frac{1}{k_0-1} \right) + \frac{r}{k_0-1} v_N(k_0-1) = v_N(k_0-1). \]

It shows that \( v_N = v_N(k_0-1) \) and the stopping rule \( \tau^* = \min\{1 \leq r \leq N - 1 : r \geq k_0, R_r = 1\} \lor N \) is optimal.

\[ \begin{array}{|c|c|c|c|}
\hline
N & \text{Cost of decision} & \text{Cost of decision} & \text{Cost of decision} \\
\hline
5 & 2 \frac{13}{20} \cong 0.65 & 2 & 2 \frac{7}{10} \cong 0.66667 \\
10 & 3 \frac{11}{12} \cong 0.73333 & 3 & 4 \frac{5}{6} \cong 0.66667 \\
15 & 4 \frac{31}{40} \cong 0.775 & 4 & 5 \frac{5}{6} \cong 0.66667 \\
50 & 7 \frac{86}{100} \cong 0.868571 & 7 & 9 \frac{7}{8} \cong 0.875 \\
100 & 10 \frac{905}{100} \cong 0.905446 & 10 & 14 \frac{3}{4} \cong 0.875 \\
\infty & 0 & \left[ 0.00251646N \right] & 0.9 \\
\hline
\end{array} \]

Let the number of applicants be going to the infinity. When the cost \( c \) is positive the value of the problem has limit less than 1 and the asymptotic threshold is bigger than 0.

**Theorem 3.4** Let us assume that \( c \in (0, +\infty) \). We have

\[ \lim_{k_0 \to \infty} v_N = 1 - c \left( c + \frac{1}{2} \right) \alpha - \frac{c\alpha}{1 - \alpha} \log(\alpha) \]

and \( \alpha \) is the unique solution of the equation \( \log(x) = \left( 1 + \frac{1}{2c} \right) (x - 1) \) in \((0, 1)\).

**Proof.** It is a consequence of Theorem 3.3 and the observation from Remark 3.2.

**Remark 3.5** It is also natural payoff structure when at the last moment \( N \) there are no cost of decision and \( c \in [0, \frac{1}{2}) \). In this case the decision maker will hesitate longer before he accepts the candidate than in the model with cost of decision at the last moment. A numerical example is given in Table 2. The form of optimal strategy is the same. The threshold \( k^*_0 \) is different. Its limit \( \frac{k^*_0}{N} \to \beta \) fulfills the equation \( \log(x) = \frac{1}{2c} (x - 1) \).
Table 2
Optimal strategy and expected payoff when there is no cost at last moment.

| N  | Cost of decision | \( c = 0 \) | \( c = \frac{1}{10} \) | \( c = \frac{2}{10} \) |
|----|------------------|------------|----------------|----------------|
| 5  | 2                | \( \frac{13}{20} \) \( \approx 0.65 \) | 3              | 3              | 0.566667       |
| 10 | 3                | \( \frac{11}{15} \) \( \approx 0.7333 \) | 4              | 5              | 0.626485       |
| 15 | 4                | \( \frac{31}{40} \) \( \approx 0.775 \) | 5              | 6              | 0.662696       |
| 50 | 7                | 0.868571   | 9              | 14             | 0.729829       |
| 100| 10               | 0.905446   | 14             | 22             | 0.755734       |
| \( \infty \) | 1 | [0.00697715N] | 0.9            | [0.107355N] | 0.8            |

4 Final remarks

The cost of decision included in this model gives parameter to measure the fear of grader that his decision is too early. One can also imagine that the grader is able to observe the true value of the item over some fixed threshold, the level of the price acceptable by him. In this case, the value of the threshold determine the expected number of observation to the acceptance (see Porosinski and Szajowski (2000)). Such partial observation is easy to realize by human being and it is natural behaviour for many traders. They do not accept prices below some threshold.

In many real problems one can observe that the decision maker hesitates to long and postpones the final decision. He rejects relatively best option too long. It looks that he fears to loss the potential options. The level of fear can be dependent on the value of the item or independent. The model of choice for such decision maker could be based on the multicriteria optimal stopping models considered by Gnedin (1981), Ferguson (1992), Samuels and Chotlos (1986) and recently by Sakaguchi and Szajowski (2000) and Bearden et al. (2005). In this model the one variable is related to the value or rank of the applicant being searched. The second coordinate would be a measure of undefined risk related to the decision process which the decision maker is feeling. From this point of view the research is needed to adopt the proper model for the considered case of the item selection. It also open the theoretical investigation to formulate variation of the best choice selection.

References

Bartoszynski, R., Govindarajulu, Z., 1978. The secretary problem with interview cost. Sankhya, Ser. B 40, 11–28.
Bearden, J. N., 2006. A new secretary problem with rank-based selection and cardinal payoffs. J. Math. Psychology 50, 58 – 59.

Bearden, J. N., Murphy, R. O., Rapoport, A., 2005. A multi-attribute extension of the secretary problem: Theory and experiments. J. Math. Psychology 49, 410 – 422.

Cayley, A., 1875. Mathematical questions with their solutions. The Educational Times 23, 18–19.

Corbin, R. M., 1980. The secretary problem as a model of choice. J. Math. Psychol. 21, 1–29.

Dynkin, E., Yushkevich, A., 1969. Theorems and Problems on Markov Processes. Plenum, New York.

Ferguson, T., 1989. Who solved the secretary problem? Statistical Science 4, 282–296.

Ferguson, T. S., 1992. Best-choice problems with dependent criteria. In: Ferguson, T. S., Samuels, S. M. (Eds.), Strategies for Sequential Search and Selection in Real Time, Proceedings of the AMS-IMS-SIAM Joint Summer Research Conferences held June 21-27, 1990. Vol. 125 of Contemporary Mathematics. American Mathematical Society, Providence, Rhode Island, University of Massachusetts at Amherst, pp. 135–151.

Frank, A., Samuels, S., 1980. On an optimal stopping of Gusein-Zade. Stoch. Proc. Appl. 10, 299–311.

Freeman, P., 1983. The secretary problem and its extensions: a review. Int. Statist. Rev. 51, 189–206.

Gardner, M., 1960a. Mathematical games. Scientific American 202 (1), 150–156.

Gardner, M., 1960b. Mathematical games. Scientific American 202 (3), 172–182.

Gilbert, J., Mosteller, F., 1966. Recognizing the maximum of a sequence. J. Amer. Statist. Assoc. 61 (313), 35–73.

Gnedin, A., 1981. Multicriterial problem of optimum stopping of the selection process. Autom. Remote Control 42, 981–986.

Gusein-Zade, S., 1966. The problem of choice and the optimal stopping rule for a sequence of independent trials. Theory Probab. Appl. 11, 472–476.

Mori, T., 1985. Hitting a small group of a middle ranked candidates in the secretary problem. In: Proc. of the 5th Pann.Symp. on Math.Stat.,May 1985, Visegrad.

Porosinski, Z., Szajowski, K., 2000. Full-information best choice problem with random starting point. Math. Jap. 52 (1), 57–63.

Resnick, S. I., 1987. Extreme values, regular variation, and point processes. Vol. 4 of Applied Probability. Springer-Verlag, New York.

Rose, J., 1982a. Selection of nonextremal candidates from a sequence. J. Optimization Theory Appl. 38, 207–219.

Rose, J., 1982b. Twenty years of secretary problems: a survey of developments in the theory of optimal choice. Management Studies 1, 53–64.

Sakaguchi, M., Szajowski, K., 2000. Mixed-type secretary problems on sequences of bivariate random variables. Math. Jap. 51 (1), 99–111.

Samuel-Cahn, E., October 2005. When should you stop and what do you get? some secretary problems. Discussion Paper 407, Deparment of Statistics, The Hebrew University
Samuels, S., 1991. Secretary problems. In: Ghosh, B., Sen, P. (Eds.), Handbook of Sequential Analysis. Marcel Dekker, Inc., New York, Basel, Hong Kong, pp. 381–405.
Samuels, S. M., Chotlos, B., 1986. A multiple criteria optimal selection problem. In: Ryzin, J. V. (Ed.), Adaptive statistical procedures and related topics. Proceedings of the Symposium on Adaptive Statistical Procedures and Related Topics, held at Brookhaven National Laboratory, June 1985. No. 8 in IMS Lect. Notes Monogr. Ser. Institute of Mathematical Statistics, Beachwood, OH 44122, U.S.A., pp. 62–78.
Seale, D., Rapoport, A., 1997. Sequential decision making with relative ranks: An experimental investigation of the ”secretary problem”. Organizational Behaviour and Human Decision Processes 69, 221–236.
Seale, D., Rapoport, A., 2000. Optimal stopping behavior with relative ranks: The secretary problem with unknown population size. J. Behavioral Decision Making 13, 391–411.
Smith, M., 1975. A secretary problem with uncertain employment. J. Appl. Probab. 12, 620–624.
Smith, M., Deely, J., 1975. A secretary problem with finite memory. J. Amer. Stat. Assoc. 70, 357–361.
Suchwalski, A., Szajowski, K., 2002. Non standard, no information secretary problems. Sci. Math. Japonicae 56, 443 – 456.
Szajowski, K., 1982. Optimal choice problem of a-th object. Matem. Stos. 19, 51–65, in Polish.
Yang, M., 1974. Recognizing the maximum of a random sequence based on the relative rank with the backward solicitation. J. Appl. Prob. 11, 504–512.
Yeo, G. F., 1998. Interview costs in the secretary problem. Aust. N. Z. J. Stat. 40 (2), 215–219.