Adiposity and ischemic and hemorrhagic stroke
Prospective study in women and meta-analysis

ABSTRACT

Objective: To compare associations of body mass index (BMI) with ischemic stroke and hemorrhagic stroke risk, and to review the worldwide evidence.

Methods: We recruited 1.3 million previously stroke-free UK women between 1996 and 2001 (mean age 57 years [SD 5]) and followed them by record linkage for hospital admissions and deaths. We used Cox regression to estimate adjusted relative risks for ischemic and hemorrhagic (intracerebral or subarachnoid hemorrhage) stroke in relation to BMI. We conducted a meta-analysis of published findings from prospective studies on these associations.

Results: During an average follow-up of 11.7 years, there were 20,549 first strokes, of which 9,993 were specified as ischemic and 5,852 as hemorrhagic. Increased BMI was associated with an increased risk of ischemic stroke (relative risk 1.21 per 5 kg/m^2 BMI, 95% confidence interval 1.18–1.23, p < 0.0001) but a decreased risk of hemorrhagic stroke (relative risk 0.89 per 5 kg/m^2 BMI, 0.86–0.92, p < 0.0001). The BMI-associated trends for ischemic and hemorrhagic stroke were significantly different (heterogeneity: p < 0.0001) but were not significantly different for intracerebral hemorrhage (n = 2,790) and subarachnoid hemorrhage (n = 3,062) (heterogeneity: p = 0.5). Published data from prospective studies showed consistently greater BMI-associated relative risks for ischemic than hemorrhagic stroke with most evidence (prior to this study) coming from Asian populations.

Conclusions: In UK women, higher BMI is associated with increased risk of ischemic stroke but decreased risk of hemorrhagic stroke. The totality of the available published evidence suggests that BMI-associated risks are greater for ischemic than for hemorrhagic stroke.

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review and meta-analysis of published prospective studies relating BMI to risk of ischemic and hemorrhagic stroke.

**METHODS Study participants.** Between 1996 and 2001, 1.3 million middle-aged women in the United Kingdom were recruited to the Million Women Study at breast cancer screening centers in England and Scotland. Each participant returned a recruitment questionnaire on health and lifestyle characteristics, including current height and weight. In 2006–2009, direct measurements of height and weight were obtained for a subset of participants 9 years after recruitment on average. We followed the entire cohort for deaths, emigrations, and hospital admissions (as inpatients or day cases) by linkage to National Health Service (NHS) central registers and electronic hospital records. Full details of the study design, methods, survey questionnaires, and information about data sharing can be found on the study’s website (www.millionwomenstudy.org).

**Exposure assessment.** Using information from the recruitment questionnaire, study participants were grouped by BMI (reported weight in kilograms divided by the square of the reported height in meters) in 5 categories: <22.5, 22.5–<25, 25–<27.5, 27.5–<30, ≥30 kg/m². To allow for measurement error and changes in BMI over time, when estimating trends we scored each category as the mean measured BMI among the subset of women in that category for whom direct measurements were available: 22.2, 25.2, 27.6, 30.3, and 34.4 kg/m², respectively. We defined other baseline exposures as follows: region (Scotland and the 9 extant Cancer Registry regions in England), deprivation category (a measure of socioeconomic status based on quintiles of the Townsend index, an area-based index of deprivation), self-reported strenuous exercise (“enough to cause sweating or a fast heartbeat”; none, <1, 1, 2+ times per week), moderate alcohol intake (none, 0.5–3, >3+ units per week, where 1 unit is assumed to contain 10 grams of pure alcohol), smoking (never, past, current: <15, 15+ cigarettes/d), and height (<160, 160–165, 165+ centimeters). To examine the role of factors that might mediate associations of BMI with stroke, we grouped participants according to self-reported treatment for relevant conditions at recruitment: hypertension (no, yes), high blood cholesterol (no, yes), and diabetes (no, yes).

**Outcome assessment.** Women with self-reported prior stroke at recruitment, or mention of prior cerebrovascular disease (ICD-9 codes 430–438 or ICD-10 codes I60-I69) in the electronic hospital record, and those who did not report their height or weight, were excluded from all analyses. We defined the first stroke as the earliest postrecruitment hospital admission mentioning stroke, if any, or death with stroke certified as the underlying cause. Electronic hospital records were available for April 1, 1997, to March 31, 2011, in England (Hospital Episode Statistics) and January 1, 1981, to December 31, 2008, in Scotland (Scottish Morbidity Record); death and emigration records were available up to December 31, 2012. All sources provided date and causes coded to the ICD-10 for each event during the study period. Using ICD-10, we defined subarachnoid hemorrhage as I60, intracerebral hemorrhage as I61, and ischemic stroke as I63. Stroke of unspecified type (I64) includes both ischemic and hemorrhagic strokes. We used primary care data from 2 different sources to assess the numbers of strokes specified either as hemorrhagic or as ischemic in the primary care records of women whose stroke was unspecified in hospital records. Primary care data, for those women for whom this information was available, came from (1) information obtained by record linkage to the Clinical Practice Research Datalink (CPRD) (see cprd.com for further details), which is available for 8% of the cohort; and (2) that reported by primary care physicians in a postal survey requesting further information about a random subsample of 1,004 women with a hospital admission for stroke.

For the small proportion of hospital stroke admissions with more than 1 recorded primary pathologic type (0.3%), we used the type recorded last during the admission, on the assumption that diagnostic accuracy would improve during the hospital stay. A sensitivity analysis was conducted in which the endpoint was restricted to admissions with stroke as the first diagnosis listed (92% of all admissions listing stroke as a diagnosis).

**Statistical analysis.** Observation for stroke outcomes began at recruitment or the start of electronic hospital records (whichever was later) and ceased at death, emigration, or the end of electronic hospital records (whichever was earlier), with censoring at any stroke event. We estimated hazard ratios (subsequently referred to as relative risks) for stroke in each BMI category relative to a reference group (<22.5 kg/m²) by Cox regression, taking attained age as the underlying time variable, stratifying by region, and adjusting for potential confounders (deprivation, physical exercise, alcohol intake, smoking, and height). We computed group-specific 95% confidence intervals (CIs) by estimating the variance of the log risk for each group.

We also estimated log-linear trends in risk over the 5 categories of BMI, with conventional 95% CIs, scoring each category as the mean within-category measured BMI. We expressed all trends as relative risks per 5 kg/m², and assessed heterogeneity between trend estimates with a 2-sided chi² contrast test, using a 1% significance level (rather than 5%) to allow for repeated testing. Relative risks are presented separately after adjustment for age and region only, and after additional adjustment for deprivation, exercise, alcohol, smoking and height.

Sensitivity analyses were also conducted to assess the effect of excluding the first 5 years of follow-up (to assess potential effects of reverse causation, whereby preclinical disease may affect BMI), and of additionally excluding women with self-reported prior heart disease, thrombosis, diabetes, or cancer (not just prior stroke), and of adjustment for previous history of cancer. We also assessed the effect of excluding underweight women (BMI <18.5 kg/m²) from the analysis. All calculations used Stata version 13.0.

We repeated the trend analyses for subgroups defined by various characteristics of the women, including age, adjustment factors, and whether or not women self-reported treatment, at baseline, for hypertension, high cholesterol, or diabetes.

To assess the plausibility of possible mechanisms for the association of BMI with specific stroke subtypes, mean apolipoprotein B/A1 ratio was calculated within categories of BMI in a subsample of women with measured lipid levels.

**Systematic review and meta-analysis.** We conducted a systematic review of prospective studies of BMI and incidence of the main pathologic types of stroke. In March 2015, we searched PubMed for relevant articles in English, using combinations of the MeSH terms stroke (both ischemic and hemorrhagic), intracerebral hemorrhage, subarachnoid hemorrhage, and body mass index. Titles and abstracts of identified articles were initially screened for relevance. Potentially relevant articles were then assessed for eligibility using the following inclusion criteria: cohort study (or analysis of pooled cohort studies using individual data)
published after 1994, relating adult BMI to risk of incident stroke (fatal and nonfatal), including at least 500 strokes in total (to reduce scope for publication bias), and giving separate estimates for both hemorrhagic and ischemic stroke, adjusted for confounders (at least age, and sex or smoking where relevant) but not for potential mediators (hypertension, dyslipidemia, or diabetes).

Where several models were reported, we chose the model that was adjusted for the largest number of confounders, without being adjusted for potential mediators. Where studies overlapped, we excluded the smaller study. Where necessary, trends were estimated from categorical risk estimates by generalized least squares, with group-specific BMI means estimated from means, standard deviations, or percentiles of BMI, assuming a normal distribution. Where appropriate, we combined trend estimates for subarachnoid and intracerebral hemorrhage. We expressed all trend estimates as relative risks per 5 kg/m², grouped them by prespecified geographic area (Asia and Europe, North America, or Australia, because the distribution and etiologies of stroke subtypes differ between populations of Asian and European origin), and combined results by inverse-variance methods.

Standard protocol approvals, registrations, and patient consents. The Multi-Centre Research Ethics Committee for Anglia and Oxford approved the study. Each participant gave written informed consent to follow-up through medical records.

RESULTS Million Women Study. After exclusions for prior cerebrovascular disease (1%) and missing BMI (a further 5%), 1,277,129 women were included in the analysis (table 1). Of these, 15,683 (1%) were lost to follow-up during the study period, for example, through emigration, but contributed person-years under observation until the relevant end of follow-up date. Mean age at recruitment was 56.7 years (SD 5), and mean follow-up time was 11.7 years, with 15 million person-years of follow-up included in analyses for this study. Women with higher BMI tended to have lower socioeconomic status than leaner women, and were less likely to currently smoke, drink alcohol, or use hormone therapy, but more likely to be physically inactive, and much more likely to report being on treatment for hypertension, high cholesterol, or diabetes at recruitment.

During follow-up, there were 20,549 first strokes, of which 9,993 were coded as ischemic (cerebral infarction), 5,852 were coded as hemorrhagic (2,790 intracerebral and 3,062 subarachnoid hemorrhage), and 4,704 were of unspecified type. Increased BMI was associated with an increased risk of ischemic stroke. The relative risk of ischemic stroke per 5 kg/m² increase in BMI was 1.23 (1.20–1.26) after adjustment for age and region only, and 1.21 (1.18–1.23) after additional adjustment for all 5 potential confounders listed in the Methods section. By contrast, the corresponding relative risk of hemorrhagic stroke decreased with increasing BMI with an age- and region-adjusted relative risk of 0.88 (0.85–0.91), which was relatively unaffected by adjustment for other potential confounders (relative risk 0.89, 95% CI 0.86–0.92). This difference in BMI-associated trends in risk for ischemic vs hemorrhagic stroke was highly significant (figure 1, p < 0.0001). There was no significant difference between the BMI-associated trends for intracerebral and subarachnoid hemorrhage (figure 2, p = 0.5).

The results of sensitivity analyses using only hospital admissions with stroke listed as the first diagnosis code, or excluding the first 5 years of follow-up, or

| Table 1 | Baseline characteristics and details of follow-up for stroke, by body mass index |
|---------|---------------------------------------------------------------|
| **Body mass index, kg/m²** | <22.5 | 22.5–<25 | 25–<27.5 | 27.5–<30 | ≥30 | **Any** |
| **Total no. women** | 247,335 | 344,534 | 276,986 | 179,980 | 228,274 | 1,277,129 |
| **Age, y, mean (SD)** | 56.3 (4.9) | 56.6 (4.9) | 56.9 (4.9) | 57.0 (4.9) | 56.7 (4.8) | 56.7 (4.9) |
| **Most deprived fifth, %** | 16.5 | 16.3 | 18.7 | 21.4 | 26.2 | 19.4 |
| **No strenuous physical exercise, %** | 42.2 | 42.8 | 47.6 | 52.5 | 60.2 | 48.2 |
| **Drinks alcohol, %** | 78.9 | 80.7 | 78.1 | 74.5 | 67.1 | 76.5 |
| **Currently smokes, %** | 25.1 | 20.4 | 19.3 | 18.6 | 16.4 | 20.1 |
| **On menopausal hormone therapy, %** | 37.0 | 35.4 | 33.3 | 31.4 | 27.8 | 33.3 |
| **On treatment for hypertension, %** | 8.2 | 11.3 | 15.2 | 19.5 | 28.0 | 15.7 |
| **On treatment for diabetes, %** | 0.9 | 1.1 | 1.7 | 2.9 | 6.2 | 2.3 |
| **On treatment for high cholesterol, %** | 2.5 | 3.3 | 4.3 | 5.4 | 6.5 | 4.3 |
| **Previous cancer, %** | 3.3 | 3.2 | 3.3 | 3.4 | 3.4 | 3.3 |
| **Previous heart disease, %** | 3.0 | 3.3 | 4.4 | 5.5 | 7.7 | 4.5 |
| **Person-years to first stroke event (1000s)** | 2,903 | 4,062 | 3,249 | 2,101 | 2,642 | 14,957 |
| **No. women with first stroke event** | 3,860 | 4,855 | 4,237 | 3,109 | 4,488 | 20,549 |

Percentages for each characteristic exclude women with missing information for that item.
excluding women with any self-reported prior ill health (not just stroke), or with adjustment for prior history of cancer, did not differ materially from the main findings (in each case, trend estimates differed by no more than 2% for both ischemic and hemorrhagic stroke). Exclusion of 12,639 (1%) women with a BMI of $<25$ kg/m$^2$ had a negligible effect on the main findings (trend estimates differed by no more than 2% for both ischemic and hemorrhagic stroke).

The association between BMI and unspecified stroke (I64; $n = 4,704$) resembled the pattern of risk seen for any stroke except that the relative risk of unspecified stroke was slightly higher in the most obese category (figure e-1 at Neurology.org). This is likely to reflect a greater proportion of ischemic strokes among those of unspecified type. Additional information collected from primary care about women with a stroke in this cohort shows that ischemic strokes are overrepresented in the "unspecified stroke" category of hospital data. For women who were coded as having an unspecified stroke (I64) in their hospital records, but whose stroke was specified as either ischemic or hemorrhagic in their primary care records, the proportion of strokes specified as ischemic were 75% (54/72) based on linked primary care data (CPRD) and 89% (78/88) based on information provided directly by the patient’s primary care physician. Both these percentages are greater than the corresponding proportion of 63% ischemic stroke (9,993/15,845), based on hospital records with specified type.

There was some heterogeneity in the relative risks per $5\text{ kg/m}^2$ across subgroups of women (figure 3, $p < 0.0001$ and $p = 0.003$ for global heterogeneity among the trends for ischemic stroke and hemorrhagic stroke, respectively). For hemorrhagic stroke, the decreasing trend in relative risk with higher BMI was stronger in older women, and in those who reported at baseline being treated for hypertension or for high cholesterol. For ischemic stroke, the increasing trend in relative risk with higher BMI was weaker in physically inactive women, smokers, and those who reported at baseline being treated for hypertension. The pattern of an increasing trend in ischemic stroke and a decreasing trend in hemorrhagic stroke was evident regardless of reported treatment for high blood pressure or high cholesterol.

Among 15,739 women with information on blood lipid levels, who were not receiving treatment for high cholesterol at recruitment, mean apolipoprotein B/A1 ratio increased with increasing BMI. The mean ratios in women with BMIs (in $\text{kg/m}^2$) of $<25$, $25<30$, and $30$+ were $0.62$ (0.61–0.62), $0.67$ (0.66–0.67), and $0.68$ (0.67–0.68), respectively.

**Systematic review.** Electronic searching retrieved 495 studies. Screening titles and abstracts eliminated 388 studies. Assessing full-text articles eliminated a further 95 studies that did not fulfil one or more specific inclusion criteria. This left 12 studies, 5 from Europe, North America, and Australia,15–19 and 7 from Asia20–26, 3 of these studies reported separate estimates for men and women (table e-1).16,20,24 Including our study, type-specific trend estimates were based on 52,216 stroke events in total: 30,553 from Asia and 21,656 from Europe, North America, and Australia. For Europe, North America, and Australia, previous studies contributed only 5,811 cases; our study contributed an additional 15,845 cases.

In all but 319,25,24 of the 12 published studies included in the meta-analysis, the trend estimate was lower for hemorrhagic than ischemic stroke. For both stroke types, the pooled trend estimates were lower in
European/North American populations than in Asian populations, and within these groups there was some heterogeneity across studies (figure 4).

**DISCUSSION** In this prospective study of 1.3 million UK women, among whom 20,549 incident strokes had occurred, increased BMI was found to be associated with an increased risk of ischemic stroke but a decreased risk of hemorrhagic stroke. The magnitude of the reduction in risk associated with increasing BMI was similar for intracerebral and subarachnoid hemorrhage. Published findings from other prospective studies are broadly consistent with ours. The BMI-associated relative risk was greater for ischemic than for hemorrhagic stroke overall, and separately, in 9 of the 12 previously published studies that were included in the meta-analysis. Findings from 5 other potentially eligible studies that did not publish results in a suitable form for inclusion in the meta-analysis did not appear to contradict the findings of the meta-analysis. The difference between the trends for the 2 stroke types was greater in the European, North American and Australian studies than in the Asian studies, perhaps reflecting variation in the epidemiology of different stroke types in these different populations. As in our study, there was no evidence that the BMI association differed for subarachnoid and intracerebral hemorrhage in the 2 large Asian cohort studies that estimated separate trends for these types.

The reasons for a difference between the BMI-associated risks of hemorrhagic and ischemic stroke are unclear. Obesity is associated with well-established risk factors for stroke in general, including hypertension, dyslipidemia, and diabetes. Hypertension increases the risk of both stroke types, but some evidence suggests that dyslipidemia and diabetes specifically increase the risk of ischemic stroke. It has also been suggested that hemorrhagic stroke risk increases with decreasing serum cholesterol levels. Although we were unable to assess directly whether differences in lipid levels could account for the observed associations between BMI and stroke subtypes, the increase in apolipoprotein B/A1 ratio with increasing BMI shown here is consistent with the hypothesis that BMI increases ischemic stroke risk, and decreases hemorrhagic stroke risk, through altered lipid levels.

We followed 1.3 million women for 11.7 years on average, using routinely collected NHS data and electronic hospital records to ascertain both fatal and nonfatal strokes with virtually no loss to follow-up. There is evidence from this and other studies that diagnoses of ischemic and hemorrhagic strokes recorded in hospital records are sufficiently reliable for epidemiologic investigation. Measurement error and changes in BMI over time were minimized by the use of measured height and weight from sampled participants 9 years after recruitment to score BMI categories for trend estimation. Comparisons with primary care data, and published data on the relative frequency of pathologic types of stroke in populations of European origin, suggest that relatively few women coded as having had an unspecified stroke in the hospital data had had a hemorrhagic stroke. Thus incomplete information on stroke types in hospital data is unlikely to have materially biased the association of increased BMI with decreased risk of hemorrhagic stroke observed here. Nor are reverse causation or confounding likely to explain the results, since trend estimates changed very little after excluding the first 5 years of follow-up, and the direction of the trends differed between the 2 types of stroke. Our analyses are based on strokes occurring in women over the

Figure 2

**Relative risks of intracerebral and subarachnoid hemorrhage by BMI**

| Body mass index (category) | <22.5 | 22.5–<25.0 | 25.0–<27.5 | 27.5–<30.0 | ≥30.0 |
|---------------------------|------|-----------|-----------|-----------|------|
| Mean measured BMI within each category 9 years later | 22.2 | 25.2 | 27.6 | 30.3 | 34.4 |

| Intracerebral | Events | 639 | 738 | 571 | 379 | 463 |
|---------------|--------|-----|-----|-----|-----|-----|
| Relative risk (95% CI) | 1.00 | 0.92-1.08 | 0.77-0.89 | 0.71-0.83 | 0.77-0.85 | 0.69-0.84 |

| Subarachnoid | Events | 725 | 845 | 644 | 401 | 447 |
|--------------|--------|-----|-----|-----|-----|-----|
| Relative risk (95% CI) | 1.00 | 0.89-0.95 | 0.84-0.90 | 0.81-0.89 | 0.73-0.89 | 0.66-0.80 |

Estimates are adjusted for age, region, deprivation, physical exercise, alcohol intake, and smoking. Relative risk for each category is plotted against the mean measured BMI in that category. The size of each square is proportional to the amount of statistical information contained. BMI = body mass index; CI = confidence interval. Published with permission from Adrian Goodill.

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age of 50 that resulted in hospital admission or death. Since hemorrhagic strokes are more likely than ischemic strokes to result in hospital admission or death,\textsuperscript{37,38} this is likely to have led to an overrepresentation of hemorrhagic strokes in this study compared with other studies where stroke outcome was ascertained differently. This should not, however, bias our estimates of the association between BMI and specific types of stroke and, while we cannot necessarily assume that the findings observed here apply equally to strokes managed without admission to hospital, our findings are unlikely to

### Table: Relative Risks of Hemorrhagic and Ischemic Stroke

| Age at risk (years) | Hemorrhagic stroke | Ischemic stroke |
|---------------------|--------------------|-----------------|
| <65                 | 3.182 (0.89-0.98)  | 3.419 (1.19-1.29) |
| 65+                 | 2.670 (0.80-0.88)  | 6.574 (1.15-1.22) |
| Heterogeneity       | p=0.002            |                 |

### Baseline characteristics

| Deprivation (fifths) | Hemorrhagic stroke | Ischemic stroke |
|---------------------|--------------------|-----------------|
| 1 = least           | 1.021 (0.76-0.90)  | 1.506 (1.12-1.28) |
| 2                   | 1.068 (0.84-0.99)  | 1.747 (1.15-1.29) |
| 3                   | 1.120 (0.80-0.93)  | 1.784 (1.13-1.26) |
| 4                   | 1.222 (0.83-0.96)  | 2.119 (1.14-1.27) |
| 5 = most            | 1.376 (0.88-1.00)  | 2.779 (1.16-1.27) |
| Heterogeneity       | p=0.01             |                 |

### Exercise (strenuous)

| Exercise          | Hemorrhagic stroke | Ischemic stroke |
|-------------------|--------------------|-----------------|
| Some              | 2.507 (0.86-0.95)  | 3.851 (1.22-1.32) |
| None              | 3.101 (0.84-0.91)  | 5.595 (1.13-1.21) |
| Heterogeneity     | p=0.4              |                 |

### Alcohol

| Alcohol            | Hemorrhagic stroke | Ischemic stroke |
|--------------------|--------------------|-----------------|
| None               | 1.618 (0.83-0.94)  | 3.254 (1.14-1.24) |
| <3 units/week      | 1.850 (0.84-0.94)  | 3.238 (1.16-1.26) |
| 3+ units/week      | 2.330 (0.86-0.96)  | 3.393 (1.15-1.26) |
| Heterogeneity      | p=0.8              |                 |

### Smoking

| Smoking            | Hemorrhagic stroke | Ischemic stroke |
|--------------------|--------------------|-----------------|
| Never              | 2.086 (0.84-0.94)  | 3.895 (1.22-1.32) |
| Past               | 1.437 (0.89-1.01)  | 2.493 (1.15-1.27) |
| Current            | 2.082 (0.79-0.89)  | 2.997 (1.06-1.16) |
| Heterogeneity      | p=0.02             |                 |

### Menopausal hormone therapy

| Menopausal hormone therapy | Hemorrhagic stroke | Ischemic stroke |
|----------------------------|--------------------|-----------------|
| Never                      | 2.952 (0.85-0.93)  | 5.343 (1.17-1.24) |
| Past                       | 9.63 (0.77-0.90)   | 1.582 (1.11-1.26) |
| Current                    | 1.839 (0.88-0.99)  | 2.883 (1.18-1.30) |
| Heterogeneity              | p=0.09             |                 |

### Treatment for hypertension

| Treatment for hypertension | Hemorrhagic stroke | Ischemic stroke |
|----------------------------|--------------------|-----------------|
| No                         | 4.601 (0.85-0.92)  | 6.936 (1.13-1.20) |
| Yes                        | 1.248 (0.74-0.84)  | 3.055 (1.03-1.12) |
| Heterogeneity              | p=0.004            |                 |

### Treatment for high cholesterol

| Treatment for high cholesterol | Hemorrhagic stroke | Ischemic stroke |
|-------------------------------|--------------------|-----------------|
| No                            | 4.827 (0.87-0.93)  | 8.026 (1.16-1.22) |
| Yes                           | 272 (0.60-0.82)    | 746 (1.12-1.34)  |
| Heterogeneity                 | p=0.002            |                 |

### Treatment for diabetes

| Treatment for diabetes | Hemorrhagic stroke | Ischemic stroke |
|------------------------|--------------------|-----------------|
| No                     | 5.692 (0.85-0.91)  | 9.211 (1.14-1.20) |
| Yes                    | 156 (0.90-1.35)    | 778 (0.98-1.17)  |
| Heterogeneity          | p=0.03             |                 |

### All women

|                      | Hemorrhagic stroke | Ischemic stroke |
|----------------------|--------------------|-----------------|
| All women            | 5.852 (0.86-0.92)  | 9.993 (1.18-1.23) |

Estimates are adjusted for age, region, deprivation, physical exercise, alcohol intake, and smoking, where appropriate. Person-years are classified by age at risk or baseline characteristics of participants. Dotted lines represent the respective relative risks per 5 kg/m\(^2\) in all women. The size of each square is proportional to the amount of statistical information contained. CI = confidence interval. Published with permission from Adrian Goodill.
have been materially affected by omission of this subgroup of stroke cases.

The main finding from this study is that greater adiposity is associated with an increase in ischemic stroke, but not necessarily hemorrhagic stroke. Our results suggest qualitatively different BMI-associated risks for these 2 pathologic types of stroke in UK women, and a decreased risk of both subarachnoid and intracerebral hemorrhage with higher BMI.

AUTHOR CONTRIBUTIONS
Mary Kroll analyzed the data and drafted the manuscript. Jane Green conceived and designed the study, advised on analysis and interpretation of the data, and revised the manuscript. Valerie Beral conceived and designed the study, advised on analysis and interpretation of the data, and revised the manuscript. Cathie Sudlow advised on analysis and interpretation of the data and revised the manuscript. Anna Brown linked and analyzed data from the Clinical Practice Research Datalink (CPRD). Oksana Kirichek linked and analyzed data from the Clinical Practice Research Datalink (CPRD). Alison Price revised the manuscript. Owen Yang conducted the systematic review and revised the manuscript. Gillian Reeves conceived and designed the study, advised on analysis and interpretation of the data, revised the manuscript, and is the study guarantor.

ACKNOWLEDGMENT
The authors thank the women recruited to the Million Women Study; the staff of collaborating NHS Breast Screening Centres and General Practices; and Ben Cairns (statistical advice), Adrian Goodill (drawing the figures), and Eli Harriss (literature searching). This study is based in part on data from the Clinical Practice Research Datalink obtained under license from the UK Medicines and Healthcare products Regulatory Agency. The interpretation and conclusions contained in this study are those of the authors alone.

STUDY FUNDING
This study was funded by the UK Medical Research Council and Cancer Research UK.

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DISCLOSURE
The authors report no disclosures relevant to the manuscript. Go to Neurology.org for full disclosures.

Received November 4, 2015. Accepted in final form May 13, 2016.

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