FUNCTIONAL CENTRAL LIMIT THEOREMS FOR WIGNER MATRICES

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Abstract. We consider the fluctuations of regular functions $f$ of a Wigner matrix $W$ viewed as an entire matrix $f(W)$. Going beyond the well studied tracial mode, $\text{Tr} f(W)$, which is equivalent to the customary linear statistics of eigenvalues, we show that $\text{Tr} f(W)A$ is asymptotically normal for any non-trivial bounded deterministic matrix $A$. We identify three different and asymptotically independent modes of this fluctuation, corresponding to the tracial part, the traceless diagonal part and the off-diagonal part of $f(W)$ in the entire mesoscopic regime, where we find that the off-diagonal modes fluctuate on a much smaller scale than the tracial mode. As a main motivation to study CLT in such generality on small mesoscopic scales, we determine the fluctuations in the Eigenstate Thermalization Hypothesis [1], i.e. prove that the eigenfunction overlaps with any deterministic matrix are asymptotically Gaussian after a small spectral averaging. Finally, in the macroscopic regime our result also generalises [2] to complex $W$ and to all crossover ensembles in between. The main technical inputs are the recent multi-resolvent local laws with traceless deterministic matrices from the companion paper [3].

1. Introduction

The eigenvalues $\{\lambda_i\}_{i=1}^N$ of large $N \times N$ Hermitian random matrices $W$ form a strongly correlated system of random points on the real line. One manifestation of this feature is that their linear statistics, $\text{Tr} f(W) = \sum_{i=1}^N f(\lambda_i)$ with a regular test function $f : \mathbb{R} \to \mathbb{R}$ has a variance of order one, in fact it satisfies a central limit theorem (CLT) but without the customary $N^{-1/2}$ scaling factor. Note that Gaussian fluctuations normally emerge with the $N^{-1/2}$ factor as a cumulative effect of $N$ independent or weakly dependent random variables. Thus it is quite remarkable that CLT holds for the strongly correlated eigenvalues and the anomalous scaling alone offsets all effects of these correlations, rendering the fluctuations of $\sum_i f(\lambda_i)$ still Gaussian.

What about the fluctuations of $f(W)$ viewed as a matrix and not just considering its trace? In this paper we show that $f(W)$ tested against any bounded deterministic matrix $A$, $\|A\| \leq 1$, is still asymptotically normal, provided that $\text{Tr} AA^* \gtrsim N^\epsilon$. Our result holds in the macroscopic and in the entire mesoscopic regime, including spectral edges. More precisely, we consider the centred functional linear statistics

$$L_N(f, A) := \text{Tr} \left[ f(W)A \right] - \mathbb{E} \text{Tr} \left[ f(W)A \right] = \sum_{i=1}^N f(\lambda_i) \langle u_i, Au_i \rangle - \mathbb{E} \left[ \ldots \right],$$

(1)
where \( \mathbf{u}_i \) is the normalized eigenvector of \( W \) corresponding to \( \lambda_i \). The statistics is called macroscopic if \( f \) is \( N \)-independent and mesoscopic on scale \( N^{-\alpha} \) with some exponent \( \alpha \in (0, 1) \) if \( f \) is of the form \( f(x) = g(N^\alpha (x - E)) \) with some \( N \)-independent compactly supported function \( g \), i.e. if \( f \) lives on a scale \( N^{-\alpha} \) around a fixed energy \( E \in [-2, 2] \) in the spectrum.

One prominent motivation to study functional CLT on small mesoscopic scales is to understand the fluctuation in the Eigenstate Thermalization Hypothesis in physics [17], also known as the strong Quantum Unique Ergodicity (QUE) in mathematics [48], see [35] for further references. QUE for Wigner matrices asserts that a law of large numbers holds for the eigenvector overlaps with deterministic matrices \( A \), i.e. that \( \langle \mathbf{u}_i, A \mathbf{u}_i \rangle \) converges to the normalized trace of \( A \) as \( N \to \infty \). In our companion paper [35] we established the optimal convergence rate of order \( N^{-1/2+\epsilon} \), for any \( \epsilon > 0 \), with a very high probability. In Theorem 2.3 of the current paper we prove that the overlaps \( \langle \mathbf{u}_i, A \mathbf{u}_i \rangle \) are asymptotically Gaussian after a small spectral averaging in the index \( i \), which corresponds to the mesoscopic functional CLT for \( \mathbf{u}_i \) when \( f \) is a characteristic function supported on a small spectral interval containing about \( N^\epsilon \) eigenvalues for any arbitrary small \( \epsilon > 0 \). We remark that the Gaussian fluctuation of \( \langle \mathbf{u}_i, A \mathbf{u}_i \rangle \) is expected to hold for each \( i \) individually, but this result has only been proven for finite rank \( A \) using the Dyson Brownian motion for eigenvectors, see [9, 10, 42].

For \( A = I \), the quantity \( L_N(f, I) \) is the standard linear statistics of the eigenvalues that have been studied extensively both in the macroscopic regime by many authors [31, 32, 53, 2, 39, 50, 54, 4, 8, 26] and in the entire mesoscopic regime \( \alpha \in (0, 1) \) by He and Knowles [29, 27, 28], see also [39, 18, 13, 6, 34, 30, 1, 36, 5, 35] for related models on mesoscopic scales and [11, 12, 19, 20] for previous works on non-optimal intermediate scales. It is therefore well known that \( L_N(f, I) \) is asymptotically normal, i.e. without a further \( N^{-\alpha/2} \) normalization it satisfies a central limit theorem with a variance given by essentially the \( H^{1/2} \)-norm of \( f \), see (17). Note that the entire analysis of the special case \( A = I \) is trivial, it relies only on the eigenvalues of \( W \) and is insensitive to its eigenvectors.

For the case of general observables, we decompose \( A \) as

\[
A = \langle A \rangle I + \hat{A}_d + A_{\text{vol}}, \quad \langle A \rangle := \frac{1}{N} \text{Tr} A,
\]

where \( \hat{A}_d = A_d - \langle A_d \rangle \) is the traceless component of the diagonal part \( A_d \) of \( A \) and \( A_{\text{vol}} := A - A_d \) is the off-diagonal part of \( A \). Following this decomposition, \( L_N(f, A) \) has three different, mutually asymptotically independent Gaussian fluctuation modes, their expectations and variances are given in Theorem 2.4. On the macroscopic scale and for real symmetric Wigner matrices this result was essentially obtained by Lytova in [37]. In Theorem 2.4 we extend [37] to complex Hermitian Wigner matrices including all crossover ensembles, i.e. following the dependence on the real parameter \( \sigma := N \text{ E } w_{12} \) in its entire range \( \sigma \in [-1, 1] \) under the standard normalization \( \text{E}[w_{12}]^2 = \frac{1}{N}, \text{E}[w_{12}] = 0 \) for the off-diagonal matrix elements of \( W \).

Our main contribution, however, is to establish a similar decomposition of fluctuations for the entire mesoscopic regime, \( \alpha \in (0, 1) \), since our Theorem 2.4 also allows for mesoscopic test functions. The corresponding limiting variances are computed in Propositions 2.9–2.10. For mesoscopic test functions \( f \) the current paper contains the first results on the limiting distribution of \( \text{Tr}[f(W)A] \), with \( A \neq I \), that we extend to the two traceless modes fluctuate on a scale of order \( N^{-\alpha/2} \) in the bulk and \( N^{-3\alpha/4} \) at the edge in contrast to the \( O(1) \) fluctuation scale of \( L_N(f, I) \). Hence not only need to explore the genuine off-diagonal fluctuations involving eigenvectors, but we also need to work at a much higher accuracy to detect the relevant fluctuations that are subleading compared with the previously explored regimes. This is a major new complication not present in the \( \alpha = 0 \) macroscopic scale in [37]. Furthermore, we also show that mesoscopic linear statistics living on different scales are asymptotically independent (Theorem 2.13).

We explain the phenomenon of different fluctuation scales on the standard example of the resolvents, \( G = G(z) = (W - z)^{-1} \) with spectral parameter \( z \in \mathbb{C} \setminus \mathbb{R} \), that can be viewed as a function \( f \) of \( W \) living on scale \( \eta := 3\varepsilon > 0 \) around the point \( E := \Re z \). To understand \( \langle GA \rangle \) for a deterministic matrix \( A \), we decompose \( A \) into its tracial and traceless parts as \( A := \langle A \rangle + \hat{A} \) and write

\[
\langle GA \rangle = m\langle A \rangle + \langle A \rangle\langle G - m \rangle + \langle G\hat{A} \rangle,
\]

where \( m, \hat{A} \) are the mean vector and the mean tracial matrix of \( A \), respectively.
where $m = m(z)$ is the Stieltjes transform of the semicircle law. The first term is deterministic, the second one is asymptotically Gaussian on scale $(G - m)(z) \sim (N\eta)^{-1}$ by [28]. We prove that the last term in (3) is also Gaussian, independent of the first one, and it has size $(GA) \sim (AA^*)^{1/2}/(N\eta^{1/2})$, provided that $(AA^*) \gg (N\eta)^{-1}$. In fact, it can be further split into a diagonal and off-diagonal part following (2). Thus the fluctuation of the tracial part is much bigger than that of the traceless part in the small $\eta$ regime, however, the latter determines the fluctuation of $(GA)$ for traceless observables $(A) = 0$.

We now mention a few related works on general Gaussian fluctuations in Wigner matrices. In contrast to the extensively studied linear eigenvalue statistics, this question received much less attention in the random matrix community, although a Wigner matrix contains many other physically or mathematically relevant random modes and most of them are expected to be Gaussian (notable exception is the eigenvalue gaps that follow the Wigner-Dyson statistics). Besides Lytova’s work [7], tracial CLTs for certain minors were obtained in [24]. Special functional CLTs have been proven for Haar distributed matrices [52, 53], and for partial traces of invariant ensembles [44]. The free probability community has systematically studied Gaussian fluctuations of traces of products of a Wigner matrix and deterministic matrices via the concept of second order freeness [16, 43]. This theory has recently been extended to polynomials in several independent Wigner matrices [4, Theorems 3–4]. However, these results rely on the moment method and handle only polynomials of Wigner matrices. It is yet unclear if the moment approach can be extended to general functions on the macroscopic scale; mesoscopic scales seem inaccessible.

Finally, we mention that the fluctuation of certain specific observables may be non-Gaussian. For example, the fluctuation of matrix entries $f(W)_{ij}$ of $f(W)$ for regular test functions $f$ is a linear combination of $w_{ij}$ and an independent Gaussian of size $N^{-1/2}$, see [39, 46, 45, 23, 40]. In contrast, our result shows that $\text{Tr } f(W)A$ is always asymptotically Gaussian whenever $\|A\| \sim 1$ and $\langle AA^* \rangle \gtrsim N^{-1+\epsilon}$. Hence, the non-Gaussian components of $f(W)$ are only visible for very low rank observables $A$.

The paper is structured as follows. After this introduction, we present the main results in the next Section 2. We start with our motivating Theorem 2.3 on the Gaussian fluctuation of the overlaps $(u_i, Au_i)$ after a small spectral averaging in $i$. Then we formulate our functional CLT (Theorem 2.4) in full generality in the bulk and at the edge of the spectrum of $W$, from the macroscopic scale down to the smallest possible mesoscopic scale just above the local eigenvalue spacing. Our formulation exhibits the three distinguished fluctuation modes with their own scaling factors. Simplified formulas in the mesoscopic regime for the expectations and the variances of the limit Gaussian processes are given in Proposition 2.9 in the bulk and in Proposition 2.10, respectively. We also include all the additional effects of the fourth cumulant $\kappa_4 = N^2 \mathbb{E}|w_{12}|^4 - 2 - \sigma^4$ of the off-diagonal matrix element $w_{12}$, the parameter $\sigma = N\mathbb{E}w_{12}^2$ describing the crossover regime between complex and real symmetry class and the size of the diagonal element $w_2 = N\mathbb{E}w_{11}^2$. These three parameters appear in the exact form of the limiting expectations and variances of the three different modes of $L_N(f, A)$. Some earlier works assumed special values of these parameters, e.g. $\sigma = 0, 1$ and $w_2 = 1 + \sigma$ is a typical choice in certain more restricted definition of the Wigner ensemble. Consequently, some explicit terms did not always appear. We also identify the cases when some of these three limiting modes have vanishing variance and explain their algebraic origin in Appendix A. Finally, in Theorem 2.13 we show that fluctuations on different scales are asymptotically independent. In Section 3 we present the necessary multi-resolvent local laws: some of them have already been proven in [35], some others, especially the ones involving three resolvents, are shown here with some proofs deferred to Appendix D.2. The main technical input for all these cases is [5, Theorem 5] and its slight extension in Theorem 3.5, proven in Appendix D.1, that control the most critical fluctuation term (the so-called renormalized “underlined” term) in the self-consistent equation for products of resolvents and deterministic matrices. Some additional technical estimates are deferred to Appendix B. In Section 4 we prove a general CLT for resolvents; this section is the technical centrepiece of the current paper. Finally, in Section 5 we convert the resolvents into general functions by using Helffer-Sjöstrand type-formulas and thus prove our general functional CLT’s.
The proof of Theorem 2.3 is given in full details in Section 5, while several technical calculations for the proof of the very general Theorem 2.4 are deferred to Appendix E.

**Notations and conventions.** We introduce some notations we use throughout the paper. For integers \( k \in \mathbb{N} \) we use the notation \([k] := \{1, \ldots, k\}\). For positive quantities \( f, g \) we write \( f \lesssim g \) and \( f \sim g \) if \( f \leq Cg \) or \( cg \leq f \leq Cg \), respectively, for some constants \( c, C > 0 \) which depend only on the moments of the matrix elements, i.e. on the constants appearing in \((\cdot)\). We denote vectors by bold-faced lower case Roman letters \( \mathbf{x}, \mathbf{y} \in \mathbb{C}^k \), for some \( k \in \mathbb{N} \). Vector and matrix norms, \( \|\mathbf{x}\| \) and \( \|A\| \), indicate the usual Euclidean norm and the corresponding induced matrix norm. For any \( N \times N \) matrix \( A \) we use the notation \( \langle A \rangle := N^{-1} \text{Tr } A \) to denote the normalized trace of \( A \). Moreover, for vectors \( \mathbf{x}, \mathbf{y} \in \mathbb{C}^N \) we define

\[
\langle \mathbf{x}, \mathbf{y} \rangle := \sum_1^N x_i y_i, \quad A_{xy} := \langle \mathbf{x}, A \mathbf{y} \rangle,
\]

with \( A \in \mathbb{C}^{N \times N} \).

We will use the concept of "with very high probability" meaning that for any fixed \( D > 0 \) the probability of the \( N \)-dependent event is bigger than \( 1 - N^{-D} \) if \( N \geq N_0(D) \). Moreover, we use the convention that \( \xi > 0 \) denotes an arbitrary small constant which is independent of \( N \).

**2. Main results**

Let \( W \) be an \( N \times N \) real or complex Wigner matrix with eigenvalues \( \lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_N \) and corresponding orthonormal eigenvectors \( \mathbf{u}_1, \ldots, \mathbf{u}_N \). The eigenvalue density profile is described by the semicircular law

\[
\rho(x) = \rho_{sc}(x) := \frac{\sqrt{A-x^2}}{2\pi}, \quad \text{for } x \in [0, A]. \tag{4}
\]

On the entries of \( W \) we formulate the following assumptions.

**Assumption 2.1.** The matrix elements \( w_{ab} \) of \( W \) are independent up to Hermitian symmetry \( w_{ab} = w_{ba}^{*} \). We assume identical distribution in the sense that \( w_{ab} \overset{d}{=} N^{-1/2} X_{ab} \), for \( a < b \), \( w_{aa} \overset{d}{=} N^{-1/2} X_d \), with \( X_d \) being a real, and \( X_{od} \) being either a real or complex random variable such that \( \mathbf{E} X_{od} = \mathbf{E} X_{dd} = 0 \), \( \mathbf{E}|X_{od}|^2 = 1 \) and \( \sigma := \mathbf{E} X_{dd} \in \mathbb{R} \). In addition, we assume the existence of the high moments of \( X_{od} \), i.e. that there exist constants \( C_p > 0 \), for any \( p \in \mathbb{N} \), such that

\[
\mathbf{E}|X_{dd}|^p + \mathbf{E}|X_{od}|^p \leq C_p. \tag{6}
\]

Notice that \( \sigma \in [-1,1] \); the case \( \sigma = 0 \) corresponds to complex Hermitian Wigner matrices with \( \mathbf{E} w_{ab}^2 = 0 \), the case \( \sigma = 1 \) corresponds to real symmetric matrices, and the case \( \sigma = -1 \) corresponds Wigner matrices \( W = D + iO \), with \( D \) being a diagonal matrix and \( O \) being skew-symmetric, i.e. \( O^t = -O \).

Finally, in order to state our results compactly, we introduce the following notation to indicate that two random vectors have asymptotically equal moments.

**Notation 2.2.** For two random vectors \( X = (X_1, \ldots, X_k) \), \( Y = (Y_1, \ldots, Y_k) \), with \( k \in \mathbb{N} \), of \( N \)-dependent random variables we define of the concept of closeness in the sense of moments and we denote it as

\[
X \equiv_{N^{-c}} Y + \mathcal{O}(N^{-c'})
\]

for some \( c > 0 \), if for any polynomial \( p(x_1, \ldots, x_k) \) it holds that

\[
\mathbf{E} p(X_1, \ldots, X_k) = \mathbf{E} p(Y_1, \ldots, Y_k) + \mathcal{O}(N^{-c' + \xi}),
\]

for any small \( \xi > 0 \), where the implicit constant in \( \mathcal{O}() \) depends on \( k, \xi \), the polynomial \( p \) and the constants in Assumption 2.1.
2.1. CLT for eigenvector overlaps. As explained in the introduction, the Gaussian fluctuation of the eigenvector overlaps $\langle u_i, Au_j \rangle$ with a deterministic matrix $A$ is a fundamental question since it describes the fluctuation in the strong Quantum Unique Ergodicity for Wigner matrices. This problem has only been solved for finite rank $A$, see [9, 10, 42]. Our first theorem establishes an averaged version of this CLT for general $A$.

Theorem 2.3 (CLT for averages of eigenvector overlaps). Let $A$ be a deterministic $N \times N$ matrix with $\|A\| \leq 1$ and let $\tilde{A} := A - \langle A \rangle$ denote its traceless part. Let $\epsilon > 0$ and $K \in \mathbb{N}$ with $N^\epsilon \leq K \leq N^{1-\epsilon}$. Then for some $\omega = \omega(\epsilon) > 0$ we have the CLT at the edge:

$$\frac{1}{\sqrt{K}} \sum_{i=N-K+1}^{N} \sqrt{N} \left( \langle u_i, Au_i \rangle - \langle A \rangle \right) \stackrel{\mathcal{L}}{\approx} \mathcal{N} \left( 0, \frac{2\sqrt{\pi}}{3} (\langle AA^* \rangle + 1(\sigma = 1)\langle A \tilde{A} \tilde{A} \rangle) \right) + O_m(N^{-\omega}).$$  \hfill (6)

Moreover, for any $\delta > 0$ and $\delta N < i_0 < (1-\delta)N$ and $\sigma > -1$ we have CLT in the bulk:

$$\frac{1}{\sqrt{2K}} \sum_{|i-i_0| \leq K} \sqrt{N} \left( \langle u_i, Au_i \rangle - \langle A \rangle \right) \stackrel{\mathcal{L}}{\approx} \mathcal{N} \left( 0, \langle AA^* \rangle + 1(\sigma = 1)\langle A \tilde{A} \tilde{A} \rangle \right) + O_m(N^{-\omega}),$$  \hfill (7)

where the implicit constant in $O_m(\cdot)$ depends on $\delta$. Finally, in case $\sigma = -1$ for any fixed $c \in (\delta, 1-\delta)$ we have a slightly different CLT in the bulk:

$$\frac{1}{\sqrt{2K}} \sum_{|i-i_0| \leq K} \sqrt{N} \left( \langle u_i, Au_i \rangle - \langle A \rangle \right) \stackrel{\mathcal{L}}{\approx} \mathcal{N} \left( 0, \langle AA^* \rangle + 1(c = 1/2)\langle A \tilde{A} \tilde{A} \rangle \right) + O_m(N^{-\omega}).$$  \hfill (8)

In the next subsection we formulate the CLT for the functional linear statistics (i) in full generality for regular test functions $f$. Theorem 2.3 is a special case of such CLT on mesoscopic scales with $f$ essentially being the characteristic function of an interval. While this sharp cut-off test function formally does not satisfy the regularity condition imposed on $f$ in Theorem 2.4 below, in Section 5 we will show how to cover this special case as well.

2.2. General functional CLT. Let $g \in H_0^2(\mathbb{R})$ be a compactly supported real valued test function, then for $0 \leq a < 1$ and $|E| \leq 2$ we define the test function rescaled to a scale $N^{-a}$ around $E$ as

$$f(x) := g \left( N^a (x - E) \right).$$  \hfill (9)

The scale $a = 0$ corresponds to the macroscopic regime. The scales $0 < a < 1$ in the bulk and $0 < a < 2/3$ at the edges, $|E| = 2$, correspond to the mesoscopic regime. Our result holds uniformly in $E$, i.e. it also covers the entire transitional regime between bulk and edge.

For deterministic $N \times N$ matrices $A$, and test functions $f: \mathbb{R} \to \mathbb{R}$ defined as in (9), we define the centred linear statistics

$$L_N(f, A) := \sum_{i=1}^{N} f(\lambda_i) \langle u_i, Au_i \rangle - \mathbb{E} \sum_{i=1}^{N} f(\lambda_i) \langle u_i, Au_i \rangle.$$  \hfill (10)

For the general CLT it is natural to decompose the space of matrices in three mutually orthogonal subspaces. We will write a general matrix $A$ as

$$A = A_d + A_{od} = \langle A \rangle I + \tilde{A}_d + A_{od}, \quad \tilde{A} := A - \langle A \rangle,$$

i.e. as the sum of a constant multiple of the identity matrix, a diagonal traceless matrix $\tilde{A}_d$, and an off-diagonal matrix $A_{od}$. Given the decomposition of $A$, the linear statistics has three modes

$$L_N(f, A) = \langle A \rangle L_N(f, I) + L_N(f, \tilde{A}_d) + L_N(f, A_{od}).$$  \hfill (11)

which we prove to be asymptotically independent Gaussians.

For sake of shorter notations we denote the expectation of a function $f$ with respect to the semicircular density and its inverse by

$$\langle f \rangle_{sc} := \int_{-2}^{2} f(x) \frac{\sqrt{4-x^2}}{2\pi} \, dx, \quad \langle f \rangle_{1/sc} := \int_{-2}^{2} \frac{f(x)}{\pi \sqrt{4-x^2}} \, dx.$$
We also define the Stieltjes transform of the semicircle law

\[ m(z) = m_{\text{sc}}(z) := \int_{-\infty}^{\infty} \frac{\rho_{\text{sc}}(x)}{x - z} \, dx, \quad z \in \mathbb{C} \setminus \mathbb{R}. \]  

We set \( \rho(z) := \frac{1}{2} |3m_{\text{sc}}(z)| \) and note that \( \rho(x + i0) = \rho_{\text{sc}}(x) \).

Finally, we introduce a few notations related to the distribution of the matrix elements of \( W \). We denote the normalised fourth cumulant of the off-diagonal entries, the expectation of the square of the off-diagonal entries, and variance of the diagonal entries of \( W \) and a certain frequently used combination of them by

\[ \kappa_4 := \mathbb{E}|\chi_{\text{od}}|^4 - 2 - \sigma^2, \quad \sigma := \mathbb{E}\chi_{\text{od}}^2, \quad w_2 := \mathbb{E}\chi_{\text{d}}^2, \quad \tilde{w}_2 := w_2 - 1 - \sigma, \]  

respectively.

We now state our main result, the functional CLT in both the macroscopic \( a = 0 \) and mesoscopic \( a > 0 \) regimes. In Theorem 2.4 we rescale the traceless diagonal linear statistics \( L_N(f, A_d) \) and the off-diagonal linear statistics \( L_N(f, A_{\text{od}}) \) in such a way the limiting processes are, to leading order, \( N \)-independent except for the explicit dependence on \( \langle |A_d|^2 \rangle \) and \( \langle A_{\text{od}} A_{\text{od}}^* \rangle \), irrespective of the scaling parameter \( a \) for test functions of the form \( f \). In Subsection 2.3 below we provide explicit formulas for the mesoscopic limits of the processes in terms of \( g \), demonstrating the \( N \)-independence of \( L_N(f, A_d), L_N(f, A_{\text{od}}) \) to leading order.

**Theorem 2.4** (Macroscopic and mesoscopic functional CLT). Let \( 0 \leq a < 1 \) and define the scaling factor for any, possibly \( N \)-dependent, \( E = E_N \in [-2, 2] \) as

\[ C_N = C_{a,E}^N := \frac{N^a}{\rho_N^{a,E}}, \quad \text{where} \quad \rho_N = \rho_E^{a,E} := \begin{cases} \rho(E + iN^{-1/2}), & a > 0 \\ 1, & a = 0. \end{cases} \]

Note that \( C_N = 1 \) for the macroscopic \( a = 0 \) case. Let \( g \in H^2_2(\mathbb{R}) \) be a compactly supported function and set \( f(x) := g(N^a(x - E)) \). Let \( A \) be a deterministic matrix with \( \|A\| \leq 1 \). Then, in the limiting regime \( C_N \ll N \), the three centred linear statistics \( \xi \) are approximately distributed (in the sense of moments)

\[ \left( L_N(f, I), \sqrt{C_N}L_N(f, A_d), \sqrt{C_N}L_N(f, A_{\text{od}}) \right) \]

\[ \equiv \left( \xi_{\text{tr}}(f), \xi_d(f, A_d), \xi_{\text{od}}(f, A_{\text{od}}) \right) + O_N \left( \sqrt{C_N} \right) \]

as three independent centred \( N \)-dependent Gaussian processes \( \xi_{\text{tr}}(f), \xi_d(f, A_d), \xi_{\text{od}}(f, A_{\text{od}}) \) whenever \( \langle |A_d|^2 \rangle, \langle A_{\text{od}} A_{\text{od}}^* \rangle \gtrsim C_N N^{-1/2} \) for some \( \epsilon > 0 \). Their variances are given by

\[ \mathbb{E}|\xi_{\text{tr}}(f)|^2 = V_{\text{tr}}^1(f) + V_{\text{tr}}^2(f, \sigma) + \frac{\kappa_4}{2} \langle (2 - x^2)f \rangle_{\text{tr}} + \frac{\tilde{w}_2}{4} \langle xf \rangle_{\text{tr}}^2 \]  

\[ \mathbb{E}|\xi_d(f, A_d)|^2 = C_N \langle |A_d|^2 \rangle \left( V_{\text{d}}^1(f) + V_{\text{d}}^2(f, \sigma) + \tilde{w}_2 \langle xf \rangle_{\text{d}}^2 + \kappa_4 \langle (x^2 - 1)f \rangle_{\text{d}}^2 \right) \]  

\[ \mathbb{E}|\xi_{\text{od}}(f, A_{\text{od}})|^2 = C_N \left( \langle A_{\text{od}} A_{\text{od}}^* \rangle V_{\text{od}}^1(f) + \langle A_{\text{od}} A_{\text{od}}^* \rangle V_{\text{od}}^2(f, \sigma) \right), \]  

\[ \text{*The Gaussians are scaled such that } \xi_{\text{tr}}(f), \xi_d(f, A_d)/\langle |A_d|^2 \rangle^{1/2}, \xi_{\text{od}}(f, A_{\text{od}})/\langle A_{\text{od}} A_{\text{od}}^* \rangle^{1/2} \text{ are of order one. The } N \text{-dependence of } C_N \text{ is exactly offset by the } N \text{-dependence of } V_{\text{d}}^1(f) \text{ etc. in the mesoscopic regime, see Section 2.5.} \]
with
\begin{align}
V_1^2(f) & := \frac{1}{4\pi^2} \int_{-\pi}^\pi \frac{(f(x) - f(y))^2}{x - y} \frac{4 - xy}{(4 - x^2)(4 - y^2)} \, dx \, dy, \\
V_1^4(f, \sigma) & := \frac{1}{4\pi^2} \int_{-\pi}^\pi \int_{-\pi}^\pi f(x) f(y) \partial_x \partial_y \log \left[ \frac{(x - \sigma y)^2 + (\sqrt{4 - x^2} + \sigma \sqrt{4 - y^2})^2}{(x - \sigma y)^2 + (\sqrt{4 - x^2} - \sigma \sqrt{4 - y^2})^2} \right] \, dx \, dy,
\end{align}
(17)
(18)

For simplicity, we formulated Theorem 2.4 for linear statistics with one test function \( f \) only. Our method, however, can handle linear combinations of test functions living on different scales since the main input of Theorem 2.4, the resolvent CLT in Theorem 4.1, allows for each involved resolvent to be evaluated at its own spectral parameter with possibly very different imaginary parts. Hence, by standard polarisation, a multivariate variant of Theorem 2.4 identifies the leading term of their fluctuation. However, Theorem 2.4 identifies the leading term of \( \mathbf{E} \text{Tr} f(W) \) to an accuracy beyond its fluctuation size. For both \( \text{Tr} f(W) A_d \) and \( \text{Tr} f(W) A_{\text{odd}} \) their expectations are much smaller than their fluctuation.

For simplicity, we formulated Theorem 2.4 for linear statistics with one test function \( f \) only. Our method, however, can handle linear combinations of test functions living on different scales since the main input of Theorem 2.4, the resolvent CLT in Theorem 4.1, allows for each involved resolvent to be evaluated at its own spectral parameter with possibly very different imaginary parts. Hence, by standard polarisation, a multivariate variant of Theorem 2.4 directly follows:

**Corollary 2.5 (Multivariate CLT).** Let \( p \in \mathbb{N}, E_1, \ldots, E_p \in [-2, 2], 0 \leq \alpha_1, \ldots, \alpha_p < 1 \), and let \( g_1, \ldots, g_p \in H_0^2(\mathbb{R}) \) be compactly supported test functions and set and \( f_i(x) := g_i(N^{\alpha_i}(x - E_i)) \). Then for deterministic matrices \( A_1, \ldots, A_p \) of bounded norms, \( \|A_i\| \lesssim 1 \) the joint linear statistics \((a)\) are

\[ E_{\text{tr}}\left( f, \sigma \right) := \left( f \left( 1 - \frac{1 - \sigma^2}{(1 + \sigma)^2 - \sigma x^2} \right) \right), \quad |\sigma| < 1, \]  
(23)
approximately distributed (in the sense of moments)
\[
\left( L_N(f, I), \sqrt{C_N^{i_{	ext{odd}}}} L_N(f, (A_i)_{\text{odd}}), \sqrt{C_N^{i_{	ext{odd}}}} L_N(f, (A_i)_{\text{odd}}) \right)_{i \in [p]}
\]
\[
\cong \left( \xi_{\text{tr}}(f), \xi_{\text{st}}(f, (A_i)_{\text{odd}}), \xi_{\text{tr}}(f, (A_i)_{\text{odd}}) \right)_{i \in [p]} + O_m\left( \frac{\max_i C_N^{i_{	ext{odd}}}}{N} \right)
\]
as centred Gaussian processes \( \xi_{\text{tr}}, \xi_{\text{st}}, \xi_{\text{od}} \) of covariances obtained from the variances in Theorem 2.4 by polariisation, uniformly in \( E_i \in [-2, 2] \). The implicit constant in the \( O(\cdot) \) error term above only depends on the model parameters in Assumption 2.1 and on \( g \) via \( \|g_r\|_{L^2} \) and \( \text{supp } g_i \), in particular it is independent of \( E_i \).

**Remark 2.6** (Alternative representation of the variances in Theorem 2.4 via Chebyshev polynomials).

By a direct computation using the geometric series we find
\[
V_1^2(f) = \sum_{k \geq 1} k\langle f t_k \rangle_{1_{\text{sec}}}^2, \quad V_2^2(f, \sigma) = \sum_{k \geq 1} k\sigma^k \langle f t_k \rangle_{1_{\text{sec}}}^2
\]
\[
V_1^2(f) = \sum_{k \geq 1} \langle f u_k \rangle_{1_{\text{sec}}}^2, \quad V_2^2(f, \sigma) = \sum_{k \geq 1} \sigma^k \langle f u_k \rangle_{1_{\text{sec}}}^2
\]
where \( t_k(x) := t_k(x/2), u_k(x) := U_k(x/2) \) and \( T_k, U_k \) are the \( k \)-th Chebyshev polynomial of the first and second kind, i.e. \( T_k(\cos \theta) = \cos(k\theta), U_k(\cos \theta) = \sin((k + 1)\theta)/\sin \theta \). In particular, we can recover the representation of \( E[|\xi_{\text{tr}}(f)|^2] \) obtained in [4, Eq. (3.5)] and write
\[
E[|\xi_{\text{tr}}(f)|^2] = \sum_{k \geq 1} k(1 + \sigma^k) \langle f t_k \rangle_{1_{\text{sec}}}^2 + 2(k_4 + \sigma^2) \langle f t_2 \rangle_{1_{\text{sec}}}^2 + w_2 \langle f t_1 \rangle_{1_{\text{sec}}}^2.
\]

Similarly, for the \( \xi_{\text{st}}, \xi_{\text{od}} \) we obtain
\[
E[|\xi_{\text{st}}(f, A_4)^2] = \langle |A_4|^2 \rangle \sum_{k \geq 3} (1 + k^4) \langle f u_k \rangle_{1_{\text{sec}}}^2 + w_2 \langle f u_1 \rangle_{1_{\text{sec}}}^2 + (k_4 + 1 + \sigma^2) \langle f u_2 \rangle_{1_{\text{sec}}}^2
\]
\[
E[|\xi_{\text{od}}(f, A_{\text{od}})|^2] = \sum_{k \geq 1} (\langle A_{\text{od}} A_{\text{od}} \rangle_{1_{\text{sec}}} + \sigma^2) \langle A_{\text{od}} A_{\text{od}} \rangle_{1_{\text{sec}}}
\]
Note, that (26)–(27) are sums of non-negative terms since \( w_2 \geq 0 \) and \( k_4 = E[|\xi_{\text{od}}|^4] - 2 - \sigma^2 \geq -1 - \sigma^2 
\]
due to \( E[|\xi_{\text{od}}|^4] \geq (E[|\xi_{\text{od}}|^2])^2 \). Similarly, (28) is a sum of non-negative terms since \( \sigma^2 \langle A_{\text{od}} A_{\text{od}} \rangle_{1_{\text{sec}}} \geq -\langle A_{\text{od}} A_{\text{od}} \rangle_{1_{\text{sec}}} \).

**Remark 2.7** (Explicit formulas for \( \sigma = \pm 1 \)). The limits of (23) are explicitly given by
\[
E_{\text{tr}}(f, 1) = \langle f \rangle_{1_{\text{sec}}} - \frac{f(2) + f(-2)}{2}, \quad E_{\text{tr}}(f, -1) = \langle f \rangle_{1_{\text{sec}}} - f(0).
\]
For the variances in case \( \sigma = 1 \) we have \( \xi_{\text{tr}}(f, 1) = V_1^2(f) \) and \( V_2^2(f, 1) = V_2^2(f) \), while for \( \sigma = -1 \) we have
\[
V_1^2(f, -1) = \frac{1}{4\pi^2} \int_{-\infty}^{\infty} \frac{f(x) - f(y)}{(x - y)^2} \frac{4 - xy}{\sqrt{(4 - x^2)(4 - y^2)}} \, dx \, dy,
\]
\[
V_2^2(f, -1) = \langle f(x) f(-x) \rangle_{1_{\text{sec}}} - \langle f \rangle_{1_{\text{sec}}}^2.
\]

**Remark 2.8** (Cases of vanishing variance in Theorem 2.4). From the Chebyshev representation in Remark 2.6 we can easily identify the necessary and sufficient conditions for the processes \( \xi_{\text{tr}}, \xi_{\text{st}}, \xi_{\text{od}} \) to vanish.

(a) \( \xi_{\text{tr}}(f) = 0 \) if and only if \( f \) is of the form
\[
f(x) = 1(\sigma = -1) \left( \phi(x) - \frac{(\phi x)_{1_{\text{sec}}}}{4} x \right) + b + 1(\sigma = 0) c x + 1(\sigma = 1) (-\sigma^2) x^2
\]
for some odd function \( \phi(-x) = -\phi(x) \) and \( b, c, d \in \mathbb{R} \).

\footnote{Note that in case \( \sigma = -1 \) the condition on \( f \) differs for the three processes. For \( \xi_{\text{od}} \) and symmetric \( A_{\text{od}} = A_{\text{od}}^* \) any odd function \( f \) results in \( \xi_{\text{od}}(f, A_{\text{od}}) = 0 \), while for \( \xi_{\text{tr}} \) and \( \xi_{\text{st}} \) only odd functions \( f \) orthogonal to \( x \mapsto x \) with respect to \( \langle \cdot \rangle_{1_{\text{sec}}} \) and \( \langle \cdot \rangle_{1_{\text{sec}}} \), respectively, result in \( \xi_{\text{tr}}, \xi_{\text{st}} \) to vanish. Thus, for example, \( \xi_{\text{od}}(x^3) = 0, \xi_{\text{tr}}(x^3 - 2x) = 0 \) and \( \xi_{\text{st}}(x^3 - 3x) = 0. \)
(b) For each fixed\(^4\) \(N\) we have \(\xi_{\hat{A}_d} = 0\) if and only if either (i) \(\hat{A}_d = 0\) or (ii) \(f\) is of the form

\[
f(x) = 1(\sigma = -1) \left( \phi(x) - \langle \phi x \rangle \right) + b + 1(u_2 = 0)cx + 1(\kappa_4 = -1 - \sigma^2)dx^2
\]

for some odd function \(\phi\) and \(b, c, d \in \mathbb{R}\).

(c) For fixed \(N\) we have \(\xi_{\hat{A}_d} = 0\) if and only if either (i) \(A_{2\hat{d}d} = 0\), or (ii) \(A_{2\hat{d}d} = -A_{2\hat{d}d}\), \(\sigma = 1\), or (iii) \(A_{2\hat{d}d} = A_{2\hat{d}d}\), \(\sigma = -1\) and \(f(x) = b + \phi(x)\) for some odd function \(\phi\).

In Appendix A we will comment on why these cases naturally yield vanishing variances.

2.3. Computation of the expectations and variances in the mesoscopic regime. Theorem 2.4 identified the expectations and the variances of the limiting processes \(\xi_{\hat{A}_d}, \xi_{\hat{A}_d}, \xi_{\hat{A}_d}\) in terms of the test function \(f\). In case of mesoscopic test functions of the form \(f(x) = g(N^\alpha(x - E))\) with some scaling exponent \(\alpha\), reference energy \(E \in [-\alpha, \alpha]\) and a compactly supported function \(g \in L^2_0(\mathbb{R})\), we may compute the leading terms of the variances in terms of \(g\). The result is different in the bulk \(|E| < 2 - \epsilon\) for any \(\epsilon > 0\) independent of \(N\) and at the edge \((E = \pm 2)\), therefore here we explicitly distinguish these two regimes. We note, however, that all error terms in our main Theorem 2.4 are valid uniformly in \(E\), so this distinction is made here only in order to obtain simple limiting formulas. The proofs of the following two propositions follows from Theorem 2.4 by simple mechanical computations, and so omitted. The variances can be conveniently expressed in terms of the \(L^2\) and \(H^{1/2}\) inner products

\[
\langle f, g \rangle_{L^2} := \int_{\mathbb{R}} f(x)g(x) \, dx, \quad \langle f, g \rangle_{H^{1/2}} := \int_{\mathbb{R}^2} \frac{f(x) - f(y)}{x - y} \frac{g(x) - g(y)}{x - y} \, dx \, dy.
\]

Proposition 2.9 (Bulk scaling asymptotics). Fix an \(\alpha \in (0, 1)\) and an \(\epsilon > 0\) (independent of \(N\)). Then for any \(|E| \leq 2 - \epsilon\), the variances and expectation in Theorem 2.4 have the following large \(N\) asymptotic behaviour

\[
V^1_{\xi}(f) &= \frac{\|g\|_{H^{1/2}}^2}{4\pi^2} + \mathcal{O}(N^{-\alpha}), \\
V^2_{\xi}(f, \sigma) &= 1(\sigma = 1) \frac{\|g\|_{H^{1/2}}^2}{4\pi^2} + 1(\sigma = -1) 1(E = 0) \frac{\langle g(x), g(-x) \rangle_{H^{1/2}}}{4\pi^2} + \mathcal{O}(N^{-\alpha}), \\
C_N V^1_{\xi}(f) &= \|g\|_{L^2}^2 + \mathcal{O}(N^{-\alpha}), \\
C_N V^2_{\xi}(f, \sigma) &= 1(\sigma = 1) \|g\|_{L^2}^2 + 1(\sigma = -1) 1(E = 0) \langle g(x), g(-x) \rangle_{L^2} + \mathcal{O}(N^{-\alpha}), \\
E_{\xi}(f, \sigma) &= 1(\sigma = -1) 1(E = 0) \frac{g(0)}{2} + \mathcal{O}(N^{-\alpha}).
\]

The implicit constants in the error terms depend only on \(\alpha, \epsilon, \|g\|_{H^{1/2}}, \|\text{ supp } g\|, \text{ and on } C_{\pi}\) from (s) and they are uniform in \(E, \sigma\) in a specific sense explained in Remark 2.11.

Proposition 2.10 (Edge scaling asymptotics). For \(E = 2\), and any \(0 < \alpha < 2/3\) the variances and expectation in Theorem 2.4 we have the scaling asymptotics

\[
V^1_{\xi}(f) = \frac{\|g(-x^2\|_{H^{1/2}}^2}{8\pi^2} + \mathcal{O}\left(N^{-\alpha/2}\right), \\
V^2_{\xi}(f, \sigma) = 1(\sigma = 1) \frac{\|g(-x^2\|_{H^{1/2}}^2}{8\pi^2} + \mathcal{O}\left(N^{-\alpha/2}\right), \\
C_N V^1_{\xi}(f) = \frac{\|g(-x^2)\|_{L^2}^2}{\pi} + \mathcal{O}\left(N^{-\alpha/2}\right), \\
C_N V^2_{\xi}(f, \sigma) = 1(\sigma = 1) \|g(-x^2)\|_{L^2}^2 + \mathcal{O}\left(N^{-\alpha/2}\right), \\
E_{\xi}(f, \sigma) = 1(\sigma = 1) \frac{g(0)}{4} + \mathcal{O}\left(N^{-\alpha/2}\right).
\]

\(^4\)Recall that the processes \(\xi_{\hat{A}_d}, \xi_{\hat{A}_d}, \xi_{\hat{A}_d}\) depend on \(N^\alpha\) through \(\hat{A}_d, A_{2\hat{d}d}\).

\(^5\)The case of the left edge, \(E = -2\) is completely analogous.
The implicit constants in the error terms depend only on $\alpha$, $\|g\|_{H^2}$, $\text{supp } g$ and on $C_p$ from (8) and on $\sigma$ in a specific sense explained in Remark 2.11.

**Remark 2.11.** Our proof also gives uniformity of the dependence on the constants $E, \sigma$ in the error terms in (33)–(34) in the following sense. In those formulas among (33)–(34) that contain $1(\sigma = 1)$, the error is uniform in $\sigma \leq 1 - \epsilon$ for any fixed $\epsilon > 0$ when $\sigma \neq 1$. Similarly, the presence of a factor $1(\sigma = -1)$ in the formula comes with uniformity for any $\sigma \geq -1 + \epsilon$ whenever $\sigma \neq -1$. Finally, in terms with $1(E = 0)$ in (33) we have uniformity for any $|E| \geq \epsilon$, whenever $E \neq 0$. In all other terms, our result is uniform for all $|E| \leq 2 - \epsilon$. See also Remark 2.15.

**Remark 2.12.** In contrast to the macroscopic scale, note that on the mesoscopic scale the limits in Propositions 2.9–2.10 are independent on $\kappa_4$ and $w_2$ and their $\sigma$-dependence is via a very simple characteristic function. This shows that the mesoscopic fluctuations are less sensitive to the details of the ensemble, in agreement with the general paradigm that more local statistics are more universal. In fact for $\sigma > -1$ the appearance of $1(\sigma = 1)$ in the variance $V_2^2, V_3^2$ corresponds to a factor of 2 difference between real symmetric and complex Hermitian symmetry classes. Furthermore, for $\sigma = -1$, assuming $w_2 = 0$, we have $W = iO$, where $O = -O^T$ is a skew symmetric matrix, in particular the spectrum of $W$ is symmetric with respect to zero, i.e. the eigenvalues around some energy $E$ and $-E$ are strongly dependent. On mesoscopic scale this feature is relevant only for $E = 0$ and it changes the expectation and the variance. In particular, for antisymmetric test functions, $g(x) = -g(-x)$, we have $L_N(f, A) = 0$, and indeed, the variances in Proposition 2.9 add up to zero in this case.

Additionally, we prove that the linear statistics for test functions living on different scales are asymptotically independent. The proof of the following theorem follows by standard arguments completely analogous to the proof of Theorem 2.4 and is presented in Section E.3.

**Theorem 2.13.** Let $\epsilon > 0$ and $E_1, E_2 \in [-2 + \epsilon, 2 - \epsilon]$, $0 \leq a_1 \neq a_2 < 1$ and let $g_1, g_2 \in H^2_\sigma(\mathbb{R})$ be compactly supported functions and set $f_i(x) := g_i(N^{a_i}(x - E_i))$. Then the limiting Gaussian processes $\xi_{11}(f_1), \xi_{11}(f_2)$ from Theorem 2.4 are asymptotically independent in the sense

$$|\text{Cov}(\xi_{11}(f_1), \xi_{11}(f_2))| \lesssim N^{-(a_1 - a_2)}.$$

(35)

Similarly, for bounded deterministic matrices $A_1, A_2$ the processes $\xi_{01}, \xi_{01}$ are asymptotically independent in the sense

$$|\text{Cov}(\xi_{01}(f_1), \xi_{01}(f_2))| + |\text{Cov}(\xi_{01}(f_1, A_1), \xi_{01}(f_2, A_2))| \lesssim N^{-(a_1 - a_2)/2}.$$

(36)

To make our presentation simpler we stated this result only in the bulk, but our proof naturally yields the independence of linear statistics living on different scales uniformly in the spectrum. Moreover, the same argument also yields independence of linear statistics living on the same scale at distant energies, i.e. for $a_1 = a_2 = a$ and $|E_1 - E_2| \gg N^{-a}$.

Theorem 2.13 together with Theorem 2.4 imply the asymptotic independence of linear statistics living on different scales $0 \leq a_1 \neq a_2 < 1$ in the sense

$$|\text{Cov}(L_N(f_1, I), L_N(f_2, I))| \lesssim N^{-(a_1 - a_2)} + N^{(a_1 - 1)/2} + N^{(a_2 - 1)/2},$$

(37)

and similarly for $\sqrt{C_N} L_N(f_1, A_0), \sqrt{C_N} L_N(f_1, A_{\text{opt}}).$ We note, however, that for large $|a_1 - a_2|$ the estimate on the covariance of linear statistics in (37) may be larger than that of the limiting processes in (33) owing to the error terms from Theorem 2.4.

### 2.4. Related earlier results and miscellaneous remarks.

The linear eigenvalue statistics $\sum_1 f(\lambda_i)$ have been extensively studied, and a CLT has been proven for macroscopic test functions as well as for mesoscopic test functions down to the optimal scale both in the bulk and at the edge, hence our results on $\xi_{11}(f)$ are not new, we only listed them for completeness. More precisely, the explicit form of the variance $E[\xi_{11}(f)]^2$ for macroscopic test functions in (4) exactly agrees with [38, Eq. (3.92)] for $w_2 = 2/\beta$ and with [50, Eq. (1.10)] for the case when $w_2 \neq 2/\beta$. Note that the parameter $\beta$, customary in random matrix theory distinguishing between the real symmetric ($\beta = 1$) and complex Hermitian ($\beta = 2$) symmetry classes, corresponds to $\beta = 2/(1 + \sigma)$ with our notation in the cases $\sigma = 0, 1$. 
For mesoscopic test functions the variance \( E[\xi_1(f)^2] \) in (44) with (33) in the bulk and with (34) at the edge exactly agree with [35, Eq. (2.22)] and [35, Eq. (2.23)], [29, Eq. (2.6)], respectively, in case of \( \sigma = 0, 1 \). Our formulas for general \( \sigma \) agree with the results in [27] for \( \sigma \in (-1, 1] \), however the final formula for the variance in case \( \sigma = -1 \) appears to be wrong in [27] (probably the error stems from [27, Eq. (6.29)] overlooking that \( \{T\} \) is not far away from zero, in fact \( |T| \sim \eta \) in this case).

As far as the expectation (density of states) is concerned, the explicit formula for \( E \sum f(\lambda_i) \) in (22) with (23) exactly agrees with the formula given in [4, Theorem 1.1] for \( \sigma \in \{0, 1\} \) and with [4, Eq. (1.4)] for the general case. We also mention that for the Gaussian case explicit \( N \)-dependent formulas are obtained in [49] on the density of states by supersymmetric methods.

The joint linear statistics of eigenvalues and eigenvectors with observable \( A \neq I \), i.e quantities \( \text{Tr}[f(W)A] = \sum_i f(\lambda_i)(u_i, Au_i) \) are much less studied. For macroscopic test functions \( f \) the variances \( E[\xi_1(f, A_d)]^2 \), \( E[\xi_{od}(f, A_{od})] \) in (44) exactly agree with [37, Eq. (4.16), Eq. (4.19)] in the real symmetric case. For mesoscopic test functions \( f \) the current paper achieves the first results on the limiting distribution of \( \text{Tr}[f(W)A] \), with \( A \neq I \), in particular, explicit formulas for \( E[\xi_1(f, A_d)]^2 \) and \( E[\xi_{od}(f, A_{od})]^2 \) in (54)–(66), with their limiting behaviour in (33) and (34), are new.

Remark 2.14. In (9) we assumed that \( g \in H^2_0(\mathbb{R}) \) to make the proof cleaner. The proof of the functional CLT (Theorem 2.4) on the macroscopic scale (\( a = 0 \)) presented in Appendix \( E \) would work exactly in the same way if \( f \in W^{2+\epsilon, 3}(\mathbb{R}) \), for some small fixed \( \epsilon > 0 \). The only difference is that throughout the proof we have to replace \( f \) by its cut-off version, \( f_\chi := f\chi \), with \( \chi \) a smooth cut-off function that is equal to one on \([-5, 5]\) and equal to zero on \([-10, 5]^c\).

Remark 2.15. The formulas in Propositions 2.9–2.10 indicate a somewhat different limiting expectation and variance when \( \sigma = \pm 1 \) in contrast to the \( |\sigma| < 1 \) case. With our methods it is also possible to study the transitional regime, where \( 1 - |\sigma| \) vanishes as an \( N \)-power, as it was done for the tracial part in [27], but we refrained from doing so in order to keep the paper more transparent.

3. Local laws for multiple resolvents

Given a Wigner matrix \( W \), we define its resolvent by \( G(z) := (W - z)^{-1} \), with \( z \in \mathbb{C} \setminus \mathbb{R} \). In this paper we consider resolvents allowing the spectral parameter \( z \) to have positive or negative imaginary part, in order to conveniently account for possible adjoints of the resolvent since \( G(z)^* = G(z) \).

In this section we prove local laws for one resolvent and for certain products of two or three resolvents that will be used as an input to prove the Central Limit Theorem for resolvents in Section 4. These local laws are stated in Propositions 3.2–3.4. Additionally, in Lemma 3.6 we present an improvement for the bound of \( (x, GAWGy)^2 \) in (46), which we need only in a second moment sense. The main inputs for the proof of these local laws are the bounds in [35, Theorem 5].

As \( N \to \infty \) the resolvent \( G \) becomes approximately deterministic (local laws). Its deterministic approximation is given by \( m(z) = m_{ac}(z) \), with \( m_{ac}(z) \) being the Stieltjes transform (12) of the semi-circular law \( \rho_{ac} \) defined in (4). In particular, \( m = m(\sigma) \) is given by the unique solution of the quadratic equation

\[
-\frac{1}{m} = z + m, \quad \Im m(z) > 0.
\]

Recall that the density \( \rho(z) \) is defined as \( \rho(z) := \pi^{-1}|\Im m(z)| \).

In order to formulate the local laws concisely we introduce the commonly used notion of stochastic domination.

Definition 3.1 (Stochastic Domination). If

\[
X = \left( X^{(N)}(u) \right)_{N \in \mathbb{N}, u \in U^{(N)}} \quad \text{and} \quad Y = \left( Y^{(N)}(u) \right)_{N \in \mathbb{N}, u \in U^{(N)}}
\]

are families of non-negative random variables indexed by \( N \), and possibly some parameter \( u \), then we say that \( X \) is stochastically dominated by \( Y \), if for all \( \epsilon, D > 0 \) we have

\[
\sup_{u \in U^{(N)}} \mathbb{P} \left[ X^{(N)}(u) > N^\epsilon Y^{(N)}(u) \right] \leq N^{-D}
\]

for large enough \( N \geq N_0(\epsilon, D) \). In this case we use the notations \( X \prec Y \) and \( X = \mathcal{O}_\prec(Y) \).
In addition to the $\mathcal{O}_\sigma(\cdot)$ notation indicating a stochastic domination in the sense of arbitrary high moments, in this proof we introduce two related new notations, $\mathcal{O}_\sigma^2(\cdot), \mathcal{O}_\sigma^4(\cdot)$, indicating domination only in first and second moment sense. More precisely, we write $X = \mathcal{O}_\sigma^2(\psi)$ and $X = \mathcal{O}_\sigma^4(\psi)$ if $\mathbb{E}[X^2] \lesssim N^\epsilon \psi^2$ and $\mathbb{E}|X|^4 \lesssim N^\epsilon \psi$, respectively, for any $\epsilon > 0$ and some deterministic $\psi$. We note that we trivially have the following product estimates

\begin{align}
X = \mathcal{O}_\sigma^2(\phi), Y = \mathcal{O}_\sigma^2(\psi) & \Rightarrow XY = \mathcal{O}_\sigma^4(\phi\psi), \\
X = \mathcal{O}_\sigma^4(\phi), Y = \mathcal{O}_\sigma^2(\psi) & \Rightarrow XY = \mathcal{O}_\sigma^4(\phi\psi), \\
X = \mathcal{O}_\sigma^2(\phi), Y = \mathcal{O}_\sigma^4(\psi) & \Rightarrow XY = \mathcal{O}_\sigma^4(\phi\psi),
\end{align}

so that by (39a), in particular, $X = \mathcal{O}_\sigma^2(\psi)$ implies $X = \mathcal{O}_\sigma^4(\psi)$.

We start with the statement of the local laws for single resolvents.

**Proposition 3.2** (Single $G$ local laws). Let $z \in \mathbb{C} \setminus \mathbb{R}$. We use the notation $\eta := |\Im z|, \rho = \rho(z), m = m_{ac}(z)$. Then for any deterministic matrix $A$ with $\|A\| \leq 1$ and $\langle A \rangle = 0$ we have the averaged local laws

$$
|\langle G - m \rangle| \lesssim \frac{1}{N\eta}, \quad |\langle GA \rangle| \lesssim \frac{\sqrt{\rho}}{N\sqrt{\eta}}.
$$

(40)

Additionally, for any deterministic vectors $x, y$ such that $\|x\| + \|y\| \lesssim 1$, we have the isotropic law

$$
|\langle x, (G - m)y \rangle| \lesssim \sqrt{\frac{\rho}{N\eta}}.
$$

(41)

The local law for $\langle GA \rangle$ in (40) is proven in [15, Theorem 3]. The averaged and isotropic law for $G - m$ have been proven in [25, 7, 33].

Next, we state averaged and isotropic local laws for products of two resolvents.

**Proposition 3.3** (Local laws for two $G$‘s). Let $z_1, z_2 \in \mathbb{C} \setminus \mathbb{R}$ and let $G_i := G(z_i)$, for $i \in \{1, 2\}$. We use the notation $\eta_i := |\Im z_i|, \rho_i = \rho(z_i), m_i = m_{ac}(z_i)$, and set $K := N\eta_i \rho^*, L := N \min(\eta_i, \rho_i)$, where $\eta_+ := \eta_1 \wedge \eta_2$ and $\rho^* = \rho_1 \vee \rho_2$. Then for any deterministic matrices $A, A'$, with $\|A\| + \|A'\| \lesssim 1$ and $\langle A \rangle = \langle A' \rangle = 0$, we have the averaged local laws\footnote{The second error term in (43) is uniform in $\sigma$ as long as $|\sigma| \leq 1 - \epsilon'$ for any fixed $\epsilon' > 0.$}

\begin{align}
\langle G_1G_2 \rangle & = \frac{m_1m_2}{1 - m_1m_2} + \mathcal{O}_\sigma \left( \frac{1}{N\eta_1\eta_2} \right), \\
\langle G_1AG_2A' \rangle & = m_1m_2 \langle AA' \rangle + \mathcal{O}_\sigma \left( \frac{\rho^*}{\sqrt{K}} \right), \\
\langle G_1G_2^* \rangle & = \frac{m_1m_2}{1 - \sigma m_1m_2} + \mathcal{O}_\sigma \left( \frac{1(\sigma = \pm 1)}{N\eta_1\eta_2} + 1(|\sigma| < 1) \left[ \frac{1}{N\eta_2^*} \wedge \frac{\rho^*}{\sqrt{K}} \right] \right), \\
\langle G_1AG_2A' \rangle & = m_1m_2 \langle AA' \rangle + \mathcal{O}_\sigma \left( \frac{\rho^*}{\sqrt{K}} \right).
\end{align}

(42-44)

We also have the following bounds

$$
|\langle G_1G_2 A \rangle| + |\langle G_1G_2^* A \rangle| = \mathcal{O}_\sigma \left( \frac{\sqrt{\rho_1 \rho_2}}{N\eta_1\eta_2} \right).
$$

(45)

Moreover, for any deterministic vectors $x, y$ such that $\|x\| + \|y\| \lesssim 1$ we have the isotropic laws

$$
|\langle x, G_1G_2 y \rangle| = \frac{m_1m_2}{1 - m_1m_2} |\langle x, y \rangle| + \mathcal{O}_\sigma \left( \frac{\sqrt{\rho}}{\sqrt{N\eta_1\eta_2^*}} \right), \\
|\langle x, G_1AG_2 y \rangle| \lesssim \sqrt{\frac{\rho^*}{\eta^*}}.
$$

(46)

where $\eta^* := \eta_1 \vee \eta_2$.

Now we state averaged laws for certain products of three resolvents.
Proposition 3.4 (Local laws for three $G$’s). Let $z_1, z_2 \in \mathbb{C} \setminus \mathbb{R}$ and let $G_i := G(z_i)$, for $i \in \{1, 2\}$. We use the notation $\eta_i := 3|z_i|, \rho_i = \rho(z_i), m_i = m_{wc}(z_i)$, and set $K := N_{\eta_1, \rho_1}, L := N_{\eta_2, \rho_2}$, where $\eta = \eta_1 \wedge \eta_2$ and $\rho^* = \rho_1 \vee \rho_2, \rho^* = \rho_1 \wedge \rho_2$. Then for any deterministic matrices $A, A'$, with $\|A\| + \|A'\| \leq 1, \langle A \rangle = \langle A' \rangle = 0$, we have the averaged local laws

$$
\langle G_1 G_2 \rangle = \frac{m_1 m_2'}{(1 - m_1 m_2')^2} + \mathcal{O}_2\left( \frac{\rho^*}{L_{\eta_1 \eta_2}} \right),
$$

$$
\langle G_1 G_2 A G_1 A' \rangle = \frac{m_1^2 m_2' \langle AA' \rangle}{1 - m_1 m_2'} + \mathcal{O}_2\left( \frac{\sqrt{\rho_1 \rho_2}}{L_{\eta_1 \eta_2}} \right),
$$

$$
\langle G_1 (G_2')^2 \rangle = \frac{m_1^2 m_2'}{(1 - \sigma m_1 m_2')^2} + \mathcal{O}_2\left( \frac{\rho_1}{L_{\eta_1 \eta_2}} \right),
$$

$$
\langle G_1 G_2 A G_1 A' \rangle = \frac{m_1^2 m_2' \langle AA' \rangle}{1 - \sigma m_1 m_2'} + \mathcal{O}_2\left( \frac{\sqrt{\rho_1 \rho_2}}{L_{\eta_1 \eta_2}} \right).
$$

Additionally, we have the following bounds

$$
|\langle G_1 G_2^T A \rangle| + |\langle G_1 G_2 \rangle| = \mathcal{O}_2\left( \frac{\sqrt{\rho_1 \rho_2}}{L_{\eta_1 \eta_2}} \right).
$$

The local laws and bounds in (42)-(46) and (47)-(51) all have the structure that the first term in the rhs. is the explicit leading term. The error term in the rhs. is smaller than the typical size of the leading term using $L \gg 1$, the fact that $|m| \sim 1/|m'| \sim \rho^{-1}$, and the bound

$$
\left| \frac{1}{1 - m_1 m_2'} \right| \leq \begin{cases} 
1/\rho^*, & \text{sgn}(3z_1) = \text{sgn}(3z_2), \\
\sqrt{\rho_1 \rho_2}/\eta_1 \eta_2, & \text{else}, 
\end{cases}
$$

which follow from elementary calculus. In the sequel we will often use these local laws in their weaker form just as an upper bound for the lhs. in terms of the upper estimate on the leading term on the rhs.

For example, (42) together with (52) implies

$$
|\langle G_1 G_2 \rangle| \lesssim \frac{\rho_1 \rho_2}{\eta_1 \eta_2},
$$

and similarly for all the other local laws.

For any given functions $f, g$ of the Wigner matrix $W$ we define the renormalisation of the product $g(W)W f(W)$ (denoted by underline) as follows:

$$
g(W)W f(W) := g(W)W f(W) - \mathbb{E}[g(W)W (\partial_W f)(W)] - \mathbb{E}[g(W)W f(W)],
$$

where $\partial_W f(W)$ denotes the directional derivative of the function $f$ in the direction $W$ at the point $W$, and $W$ is an independent copy of $W$. The definition is chosen such that it subtracts the second order term in the cumulant expansion, in particular if all entries of $W$ were Gaussian then we had $\mathbb{E}[g(W)W f(W)] = 0$. Note that the definition (53) only makes sense if it is clear to which $W$ the underline refers, i.e. it would be ambiguous if $f(W) = W$. In our applications, however, each underlined term contains exactly a single $W$ factor, and hence such ambiguities will not arise.

The key inputs for the proof of the local laws with two or three $G$’s are strong bounds for renormalised products of the form $\langle W G_1 B_1 G_2 \ldots G_l B_l \rangle$. In Theorem 5 of our companion paper [5] we proved such estimates but they are in terms $\eta_i$, the minimal of all $\eta$’s, i.e. no distinction among different $\eta$’s is made. To remedy this situation, in the following Theorem 3.5 we prove a generalization of [5, Theorem 5] which allows for the proof of the local laws for two and three $G$’s with distinguished $\eta$-dependencies as stated above. Furthermore, for a few specific terms we need a somewhat stronger bound than our general Theorem 3.5 gives, but we need them only in variance sense in contrast to the high probability bounds in Theorem 3.5. These specific bounds are listed separately in Lemma 3.6. The proof of Theorem 3.5 is presented in Appendix D.1 and the proof of Lemma 3.6 in Appendix D.2.

Theorem 3.5. Fix $\epsilon > 0$, let $l, n_1, \ldots, n_l \in \mathbb{N}$, $z_1, 1, \ldots, z_1 n_1, z_2, 1, \ldots, z_2 n_2 \in \mathbb{C} \setminus \mathbb{R}$ and for $k \in [l], j \in [n_k]$ let

$$
G_k \in \{G_{k,1} G_{k,2} \cdots G_{k,n_k}, (G_{k,1} G_{k,2} \cdots G_{k,n_k})^T \}, \quad G_{k,j} \in \{G(z_{k,j}), \mathbb{E}[G(z_{k,j})] \}
$$

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and let $B_k$ be deterministic $N \times N$ matrices, and $x$, $y$ be deterministic vectors with bounded norms $\|B_k\| \lesssim 1$, $\|x\| + \|y\| \lesssim 1$. Set
\[
L := N \min_k (\eta_i \rho_k), \quad \rho^* := \max_k \rho_k, \quad \eta_* := \min_k \eta_k,
\]
with $\eta_k := |\Im z_k|$, $\rho_k := \rho(z_k) = |\Im \zeta(z_k)|/\pi$ and assume $L \geq N^\epsilon$ and $\eta_* \lesssim 1$. Let $a, t$ denote disjoint sets of indices, $a \cap t = \emptyset$, such that for each $k \in a$ we have $\langle B_k \rangle = 0$, and for each $k \in t$ exactly one of $G_k, G_{k+1}$ is transposed, where in the averaged case and $k = l$ it is understood that $G_{l+1} = G_l$.

Then with $\eta_{\ast}$ disjoint sets of indices, a case of
\[
\langle x, G_1 B_1 G_2 B_2 \cdots G_l B_l \rangle := \frac{\rho^*}{N^\eta_*} \prod_{k \in [l]} \min_{k, i} \eta_{k, i},
\]
with $\rho_k := |\Im z_k|$, $\rho_k := \rho(z_k) = |\Im \zeta(z_k)|/\pi$ and assume $L \geq N^\epsilon$ and $\eta_* \lesssim 1$. Let $a, t$ denote disjoint sets of indices, $a \cap t = \emptyset$, such that for each $k \in a$ we have $\langle B_k \rangle = 0$, and for each $k \in t$ exactly one of $G_k, G_{k+1}$ is transposed, where in the averaged case and $k = l$ it is understood that $G_{l+1} = G_l$.

Then with $a := |a|, t := |t|$, we have the following bounds:

(i) For $a = t = \emptyset$ we have
\[
|\langle x, G_1 B_1 G_2 B_2 \cdots G_l B_l \rangle| \lesssim \frac{\rho^*}{N^\eta_*} \prod_{k \in [l]} \min_{k, i} \eta_{k, i}.
\]

(ii) For $a, t \subset [l], |a \cup t| \geq 1$ we have the bound
\[
|\langle x, G_1 B_1 G_2 B_2 \cdots G_l B_l \rangle| \lesssim \frac{\rho^*}{N^\eta_*} \prod_{k \in [l]} \min_{k, i} \eta_{k, i}.
\]

(iii) For $a, t \subset [l - 1]$ and for any $0 \leq j < l$ we have the bound
\[
|\langle x, G_1 B_1 \cdots G_j B_j W G_{j+1} B_{j+1} \cdots B_{l-1} G_l y \rangle| \lesssim \frac{\rho^*}{N^\eta_*} \prod_{k \in [l]} \min_{k, i} \eta_{k, i},
\]
where the $j = 0$ case is understood as $\langle x, W G_1 B_1 \cdots B_{l-1} G_l y \rangle$.

In case $\prod_{k \in [l]} \rho_k \lesssim (\rho^*)^{b+1}$, the bounds (53)–(57) remain valid if the rhs. are multiplied by the factor $(\rho^*)^{-b-1} \prod_{k \in [l]} \rho_k$, where $b := l$ in case of (53), $b := l - a - t$ in case of (56), and $b := l - a - t - 1$ in case of (57). Moreover, for any $\eta_* \geq 1$ we have the bounds
\[
|\langle x, G_1 B_1 G_2 B_2 \cdots G_l B_l \rangle| \lesssim \frac{1}{N^\eta_*} \prod_{k \in [l]} \min_{k, i} \eta_{k, i},
\]
\[
|\langle x, G_1 B_1 \cdots G_j B_j W G_{j+1} B_{j+1} \cdots B_{l-1} G_l y \rangle| \lesssim \frac{1}{N^l / \eta_*} \prod_{k \in [l]} \min_{k, i} \eta_{k, i}.
\]

Lemma 3.6. Let $z_1, z_2, z_T \in \mathbb{C} \setminus \mathbb{R}$ and let $G = G(z), G_1 = G(z_1)$. Then, for any fixed deterministic vectors $x, y$ and matrix $A$ with $\langle A \rangle = 0$ and $\|A\| + \|x\| + \|y\| \lesssim 1$, we have
\[
|\langle x, G A W G y \rangle| = O^2 \left( \frac{\rho}{N^{1/2} \eta} \right),
\]
and
\[
|\langle W G_1 G_2 A \rangle| + |\langle W G_1 G_2^T A \rangle| = O^2 \left( \frac{\rho^*}{N^\eta} \frac{1}{\sqrt{\eta \rho^*}} \right),
\]
\[
|\langle W G_1 G_2 A \rangle| + |\langle W G_1 G_2^T A \rangle| = O^2 \left( \frac{\rho^*}{N^\eta \sqrt{\eta \rho^*}} \right),
\]
\[
|\langle W G_1 G_2 A G_1 A \rangle| + |\langle W G_1 G_2 A G_1^T A \rangle| = O^2 \left( \frac{\rho^*}{N^{1/2} \eta \sqrt{\eta \rho^*}} \right).
\]

Notice that the bound in (59) is better by a factor $\sqrt{\rho/N \eta}$ compared to (57). The bounds (60)–(62) improve upon (56) in two aspects: First, they depend on $\rho_*$ rather than $\rho^*$, and second, in the cases including transposes the bounds distinguish different $\eta$’s (note that in Theorem 3.5 it is not allowed to have both $G, G^T$ within one $G$-block, hence the bound is purely in terms $\eta_*$).

We conclude this section with the proof of Propositions 3.3–3.4.
Proof of Proposition 3.3. The local laws for $\langle G_1 A G_2 A' \rangle$, $\langle G_1 A G_2^T A' \rangle$, in (42), (44), respectively, and the bound for $G_1 A G_2$ in (46) follow by [15, Proposition 2] together with [15, Theorem 4]. The bounds in (45) follow by exactly the same proof of [15, Eq. (22)], but using the new bound (60) instead of [15, Eq. (62)] for the underlined term. Also the local law for $\langle G_1 G_2^* \rangle$ with error term $\rho^* K^{-1/2}$ for $|\sigma| < 1$ in (43) follows by [15, Theorem 4, Proposition 2]. Hence, in order to conclude the proof of Proposition 3.3 we are left with the averaged and isotropic law for $G_1 G_2$ in (42) and (46), respectively, and with the proof of the remaining cases for the local law for $\langle G_1 G_2^T \rangle$ in (43).

We first consider the local laws that involve no transposes, then at the end of the proof of Proposition 3.3 we explain the necessary changes when the transposes are considered.

By the self consistent equation for $m$ in (58), and by $G(W - z) = I$, we have

$$G = m - mWG - m(GG + m(G - m)G).$$

As a special case of (53) we have that

$$WG = W + (G)G + \sigma \frac{\sigma^2 G}{N} + \frac{w_0}{N} \text{diag}(G),$$

where for any matrix $R$ in this section we let diag$(R)$ denote the matrix of its diagonal that was denoted by $R_d$ earlier. We recall the parameters $\sigma$, $w_0$ from (13). Then by (63) and (64), it follows that

$$G = m - mWG + \frac{m \sigma}{N} G^2 G + \frac{m w_0}{N} \text{diag}(G)G + m(G - m)G.$$  \hfill (65)

We now start writing the equation for generic products of two resolvents $G_1 G_2 G_3 G_4$, where $G_i = (W - z_i)^{-1}$ and $B_1, B_2$ are deterministic matrices. Using the equation (65) for $G_1 B_1$ and writing $G_2 = m_2 + (G_2 - m_2)$, we obtain

$$G_1 B_1 G_2 B_2 = m_1 m_2 B_1 B_2 + m_1 B_1 (G_2 - m_2) B_2 - m_1 WG_1 B_1 G_2 B_2 + m_1 (G_1 B_1 G_2) G_2 B_2$$
$$+ m_1 (G_1 - m_1) G_1 B_1 G_2 B_2 + \frac{m_1 \sigma}{N} G_1^t G_1 B_1 G_2 B_2 + \frac{m_1 \sigma}{N} (G_1 B_1 G_2)^t G_2 B_1$$
$$+ \frac{m_1 w_0}{N} \text{diag}(G_1) G_1 B_1 G_2 B_2 + \frac{m_1 w_0}{N} \text{diag}(G_1 B_1 G_2) G_2 B_2,$$

where we used that

$$WG_1 B_1 G_2 = W G_1 B_1 G_2 + \langle G_1 B_1 G_2 \rangle G_2 + \frac{\sigma}{N} \langle G_1 B_1 G_2 \rangle G_2 + \frac{w_0}{N} \text{diag}(G_1 B_1 G_2) G_2,$$  \hfill (66)

with $WG_1$ from (64). The identity in (67) follows by the definition of underline in (53).

Proof of the local law for $G_1 G_2$. We divide the proof of this local law into two cases: (i) $\exists z_1 \exists z_2 < 0$, (ii) $\exists z_1 \exists z_2 > 0$. The difference in these two cases is that in (ii) the stability factor $1 - m_1 m_2$ is bounded from below by $\rho^*$, whilst in case (i) the stability factor is bounded from below only by $\eta^*$ and so it is not affordable to invert it.

We start with $\exists z_1 \exists z_2 < 0$, in this case we can use resolvent identity and the local law $|\langle G_1 - m_1 \rangle| \prec (N \eta_1)^{-1}$ from (40):

$$\langle G_1 G_2 \rangle = \frac{G_1 - G_2}{z_1 - z_2} = \frac{m_1 - m_2}{z_1 - z_2} + O_\prec \left( \frac{1}{N \eta_1 |z_1 - z_2|} \right),$$

where we used that the self consistent equation (58) for $m_1, m_2$ in the third equality. This concludes the proof of the local law for $\langle G_1 G_2 \rangle$ when $\exists z_1 \exists z_2 < 0$. 


We now consider the case \( \exists z_1, \exists z_2 > 0 \). Choosing \( B_1 = B_2 = I \) in (66), and using \( |(G_i - m_i)| \prec (N\eta_i)^{-1} \), we find that
\[
1 - m_1 m_2 + O\left( \frac{1}{N\eta_i} \right) \left( G_1 G_2 \right) = m_1 m_2 - m_1 \langle W G_1 G_2 \rangle + m_1 \langle (G_1 - m_1) (G_2 - m_2) \rangle + \frac{m_1 \sigma}{N} \langle G_1^i G_1 G_2 \rangle + \frac{m_1 \sigma}{N} \langle (G_1 G_2)^i G_2 \rangle + \frac{m_1 w_2}{N} \langle \text{diag}(G_1) G_1 G_2 \rangle + \frac{m_1 w_2}{N} \langle \text{diag}(G_1 G_2) G_2 \rangle.
\]

Using a Schwarz inequality we readily conclude that
\[
\frac{1}{N} |\langle G_1^i G_1 G_2 \rangle| \leq \frac{1}{N} \langle G_1 G_2 \rangle \langle G_1^i G_1 G_2 \rangle \prec \frac{\rho_1}{N \eta_1 \eta_2},
\]
where we used that \( \langle G_i \rangle \prec \rho_i \), that \( \eta_i G_i \prec G_i \) by Ward identity, that \( |\langle G_1 G_2 G_1^i G_2 \rangle| \leq \|G_1 G_2\| \|G_1^i G_2\| \), and that \( \|G_i\| \prec \eta_i^{-1} \). We also prove that \( |\langle (G_1 G_2)^i G_2 \rangle| \prec \rho_2 (N\eta_1 \eta_2)^{-1} \) using exactly the same computations. Additionally, we get that
\[
\frac{1}{N} |\langle \text{diag}(G_1) G_1 G_2 \rangle| = \left| \frac{1}{N^2} \sum_{i} \langle G_{1i} G_{1i} \rangle \right| \prec \frac{\sqrt{\rho_1 \rho_2}}{\sqrt{\eta_1 \eta_2}}
\]
where we used that \( |G_{1i}| \prec 1 \), and that \( |\langle G_1 G_2 \rangle_{1i}| \prec \sqrt{\rho_1 \rho_2} / (\eta_1 \eta_2) \) by a Schwarz inequality and Ward identity. The bound for \( |\langle \text{diag}(G_1 G_2) G_2 \rangle| \) is completely analogous and so omitted. Combining (69) with (70)–(71) and using that \( |\langle G_2 - m_2 \rangle| \prec (N\eta_2)^{-1} \), we finally conclude that
\[
\langle G_1 G_2 \rangle = \frac{m_1 m_2}{1 - m_1 m_2} - \frac{m_1}{1 - m_1 m_2} \langle W G_1 G_2 \rangle + O\left( \frac{1}{N \eta_1 \eta_2} + \frac{1}{N \eta_1 \rho_i \eta_2} \right).
\]

where we used that by easy computations we have \( |1 - m_1 m_2| \geq \rho^* \) and that \( \rho^* \gg \frac{1}{N\eta_i} \), by \( K = N\eta_i \rho^* \gg 1 \), to divide through the multiplicative factor in the l.h.s. of (69). Finally, using that \( |\langle W G_1 G_2 \rangle| \prec \rho^*(N\eta_1 \eta_2)^{-1} \) by Theorem 3.5, \( |1 - m_1 m_2| \geq \rho^* \) once again, and that \( \rho^* \gg \eta_i \), \( \eta_i \eta^* = \eta_1 \eta_2 \), we conclude that
\[
\langle G_1 G_2 \rangle = \frac{m_1 m_2}{1 - m_1 m_2} + O\left( \frac{1}{\eta_1 \eta_2} \right).
\]

Proof of the local law for \( \langle x, G_1 G_2 y \rangle \). The proof of the isotropic law for \( G_1 G_2 \) is very similar to the proof of the averaged law above, hence we explain only the minor differences. Similarly to the averaged local law, the case \( \exists z_1, \exists z_2 > 0 \) trivially follows by resolvent identity. In the opposite case, choosing \( B_1 = B_2 = I \) in (66), and that \( |(G_i - m_i)| \prec (N\eta_i)^{-1} \), we find that
\[
1 + O\left( \frac{1}{N\eta_i} \right) \langle x, G_1 G_2 y \rangle = m_1 m_2 \langle x, W G_1 G_2 y \rangle + m_1 \langle G_1 G_2 \rangle \langle x, G_2 y \rangle + O\left( \frac{\rho^*}{\sqrt{N \eta_i \eta_2}} \right)
\]
where we used that the terms with a pre-factor \( \sigma \) or \( w_2 - 1 - \sigma \) can be estimated by \( N^{-1} \rho^*(N\eta_1 \eta_2)^{-1} \) using a Schwarz inequality similarly to (70)–(71). Then using that
\[
\langle G_1 G_2 \rangle = \frac{m_1 m_2}{1 - m_1 m_2} + O\left( \frac{1}{N \eta_1 \eta_2} \right),
\]
by (72), and that \( |\langle x, W G_1 G_2 y \rangle| \prec \sqrt{\rho^* (N\eta_1)^{-1/2} (\eta_i^*)^{-1}} \) by Theorem 3.5, we finally conclude that
\[
\langle x, G_1 G_2 y \rangle = \frac{m_1 m_2}{1 - m_1 m_2} \langle x, y \rangle + O\left( \frac{\sqrt{\rho^* \eta_i^*}}{\sqrt{N \eta_1 \eta_2}} + \frac{1}{N \eta_i \rho^*} \right).
\]
In order to conclude the proof of Proposition 3.3, we are left with considering transposes.

Proof of the local law for \( \langle G_1 G_2^* \rangle \). The proof of this local law is divided into three cases: (i) \( \sigma = 1 \), (ii) \( \sigma = -1 \), (iii) \( |\sigma| < 1 \). The main difference compared to the proof of \( \langle G_1 G_2 \rangle \) is that the two body stability factor is now given by \( 1 - \sigma m_1 m_2 \) instead of \( 1 - m_1 m_2 \).

For \( \sigma = 1 \) there is nothing else to prove since in this case \( W \) is real symmetric and so \( G_2^* = G_2 \).

The proof of the local law for \( |\sigma| < 1 \) is completely analogous to the proof of (72), modulo the bound for the underline term that is now given by \(|W G_1 G_2^*| < \rho^* (N \eta^2)^{-1}\), since the only thing we used in this proof is that the stability factor \( 1 - m_1 m_2 \) is bounded from below by \( \rho^* \). This is also the case for \( 1 - \sigma m_1 m_2 \) when \( |\sigma| < 1 \), since \( |1 - \sigma m_1 m_2| \geq 1 - |\sigma| \).

We are now left with the case \( \sigma = -1 \), when the stability factor is given by \( 1 + m_1 m_2 \). Note that when \( \sigma = -1 \) we can write \( W = D + iO \) with \( D \) being a diagonal matrix and \( O \) being an skew-symmetric matrix, i.e. \( O^T = -O \). If \( D = 0 \), and either \( 3z_1 3z_2 > 0 \) or \( 3z_1 3z_2 < 0 \) and \( |z_1 + z_2| \geq \eta^* \), using the notation \( R(z_i) := (iO - z_i)^{-1} \) and that \( R(z_i)^T = -R(-z_i) \), by resolvent identity we conclude

\[
\langle G(z_1)G(z_2)^* \rangle = -\langle R(z_1)R(-z_2) \rangle = -\frac{\langle R(z_1) \rangle - \langle R(-z_2) \rangle}{z_1 + z_2} = \frac{m_1 m_2}{1 + m_1 m_2} + \mathcal{O}\left( \frac{1}{N \eta_\ast |z_1 + z_2|} \right),
\]

where we used that \( m(-z_i) = -m(z_i) \), and the local law for \( R_t \), that holds even for Wigner matrices with zero diagonal. For \( 3z_1 3z_2 < 0 \) and \( |z_1 + z_2| < \eta^* \), using that \( R(z_2)^T = -R(-z_2) \) we proceed exactly as in the proof of the local law for \( \langle G_1 G_2 \rangle \) above. This gives the local law for \( \langle G_1 G_2^* \rangle \) in (43). In order to conclude the proof we are now left only with the case \( D \neq 0 \). In this case we use the following lemma whose proof is postponed to Appendix B.

**Lemma 3.7.** Fix \( \epsilon > 0 \). Let \( W = D + iO \) be a Wigner matrix with \( D \) being diagonal and \( O \) skew-symmetric. Denote \( G_1 = (W - z_1)^{-1} \) and \( R_t = (iO - z_1)^{-1} \), with \( z_1, z_2 \in \mathbb{C} \setminus \mathbb{R} \), such that \( \eta_\ast := |3z_1| \wedge |3z_2| \geq N^{-1+\epsilon} \) and \( \eta^* := |3z_1| \vee |3z_2| \), then for \( \sigma = -1 \) it holds

\[
\langle G_1 G_2^* \rangle = \langle R_1 R_2^* \rangle + \mathcal{O}\left( \frac{1}{N \eta_\ast \left[ \frac{1}{|z_1 + z_2|} \wedge \frac{1}{\eta^*} \right]} \right).
\]

Moreover, we also have

\[
\langle G_1^* G_2^* \rangle = \langle R_1^* R_2^* \rangle + \mathcal{O}\left( \frac{1}{N \eta_\ast |3z_2| \left[ \frac{1}{|z_1 + z_2|} \wedge \frac{1}{\eta^*} \right]} \right).
\]

In Lemma 3.7 we stated the result for three \( G \)’s as well (76) even if not needed for the proof of the local law of \( \langle G_1 G_2^* \rangle \). We will use (76) later in (654).

Finally, combining (75) with (74), we conclude the proof of the local law for \( \langle G_1 G_2^* \rangle \).

This concludes the proof of Proposition 3.3, modulo the proof of Lemma 3.7, which is postponed to Appendix B.

We conclude this section with the proof of the local laws for certain products of three resolvents. We will prove the estimates without transposed resolvents, the analogous results with transposes are proven in Appendix C.
Proof of Proposition 1.4. We start writing the equation for general products of three different resolvents $G_1, G_2, G_3$ and deterministic matrices $B_1, B_2, B_3$:

\[
G_1B_1G_2B_2G_3B_3 = m_1B_1G_2B_2B_3B_3 - m_1WG_1B_1G_2B_2B_3B_3 + m_1(G_1 - m_1)G_1B_1G_2B_2G_3B_3 + m_1(G_1B_1G_2B_2G_3B_3)G_3B_3
\]

\[
\begin{align*}
&\quad + \frac{m_1\sigma}{N} G_1G_1G_2B_2B_3B_3 + \frac{m_1\sigma}{N} (G_1B_1G_2)^2 G_2B_2G_3B_3 \\
&\quad + \frac{m_1\sigma}{N} (G_1B_1G_2B_2G_3)^2 G_3B_3 + \frac{m_1\sigma}{N} \text{diag}(G_1)G_1B_1G_2B_2G_3B_3 \\
&\quad + \frac{m_1\sigma}{N} \text{diag}(G_1B_1G_2B_2G_3)G_3B_3 + \frac{m_1\sigma}{N} \text{diag}(G_1B_1G_2B_2G_3)G_3B_3
\end{align*}
\]

(77)

where we used that

\[
WG_1B_1G_2B_2G_3 = \frac{\sigma}{N} (G_1B_1G_2B_2G_3)G_3B_3
\]

with $WG_1B_1G_2$ from (69). The identity in (78) follows by the definition of the renormalization (denoted by underline) in (53).

Proof of the local law for $\langle G_1G_2^2 \rangle$ in (47). Similarly to the proof of the local law for $\langle G_1G_2^2 \rangle$, the proof of the local law for $\langle G_1G_2^2 \rangle$ is divided into two cases: (i) $3z_1 \leq 3z_2 < 0$ or $3z_1 \geq 3z_2 > 0$ and $|z_1 - z_2| \geq \eta^*$

(ii) $3z_1 \leq 3z_2 > 0$ and $|z_1 - z_2| < \eta^*$, where we recall that $\eta^* := \eta_1 \wedge \eta_2$.

Similarly to (68), if either $3z_1 \leq 3z_2 < 0$ or $3z_1 \geq 3z_2 > 0$ and $|z_1 - z_2| \geq \eta^*$ we use the resolvent identity twice to get

\[
\begin{align*}
\langle G_1G_2^2 \rangle &= \langle G_2^2 \rangle - \langle G_1G_2 \rangle - \langle G_1 \rangle \langle G_2 \rangle + \mathcal{O}\left(\frac{1}{N^2|z_1 - z_2|}\right) \\
&= \frac{m_2'}{z_2 - z_1} + \frac{m_1 - m_2}{(z_2 - z_1)^2} + \mathcal{O}\left(\frac{1}{N^2|z_1 - z_2|} + \frac{1}{N^2|z_1 - z_2|^2}\right) \\
&= \frac{m_1m_2'}{(1 - m_1m_2^2)^2} + \mathcal{O}\left(\frac{\rho_s}{L\eta_1\eta_2}\right)
\end{align*}
\]

(79)

where in the first line we used the local law for $\langle G_2^2 \rangle$ in (42), and the identity $m_2' = m_2(1 - m_2^2)^{-1}$, and to go from the second to the third line we again used the equation of $m_1, m_2$. We remark that to estimate the error terms to go from the second to the third line we also used (80) below. The proof of this bound is postponed to Appendix B.

Lemma 3.8. Let $z_1, z_2 \in \mathbb{C} \setminus \mathbb{R}$ such that $|z_1 - z_2| \geq \eta^*$, then it holds

\[
\frac{1}{N\eta_1^2|z_1 - z_2|} \lesssim \frac{\rho_s}{L\eta_1\eta_2},
\]

(80)

where $\eta_1 = |3z_i|, \eta_s = \eta_1 \wedge \eta_2, \rho_s = \rho_1 \wedge \rho_2, \eta^* = \eta_1 \wedge \eta_2$.

Next we consider the last remaining case $3z_1 \geq 3z_2 > 0$ and $|z_1 - z_2| < \eta^*$. By (77) with $B_1 = B_2 = B_3 = I$ and $G_1 = G_2$, we have that

\[
\begin{align*}
(1 - m_1m_2)(G_1G_2^2) &= m_1(G_2^2) - m_1(WG_1G_2^2G_2) + m_1\langle G_1 - m_1 \rangle \langle G_2 - m_2 \rangle \langle G_1G_2^2 \rangle \\
&\quad + m_1(G_1G_2^2)(G_2^2) + \mathcal{O}\left(\frac{\rho^*}{N^2\eta_1}\right) \\
&= \frac{m_1m_2'}{1 - m_1m_2} + \mathcal{O}\left(\frac{1}{N^2\eta_1}\right)
\end{align*}
\]

(81)

where to go from the second to the third line we used the local laws for $\langle G_1G_2 \rangle$ and $\langle G_2^2 \rangle$ in (42), that $\langle |G_i - m_i| \rangle \ll (N\eta_i)^{-1}$ by the first local law in (40), and that $\langle |WG_iG_2^2| \rangle \ll \rho^*(N\eta_2\eta_1)^{-1}$.
by Theorem 3.5. We remark that to go from (77) to (80) we used that all the terms with a pre-factor \( N^{-1} \) in (77) are bounded by \( \rho^* (N \eta_1^2 \eta_2)^{-1} \). To make this clearer we show this bound for two representative terms:

\[
\frac{1}{N} \|(G_1^* G_1 G_2 G_2)\| \leq \frac{1}{N} (G_1 G_1^*)^{1/2} (G_1 G_2 G_2 G_2^* G_2^* )^{1/2} \leq \frac{\rho^*}{N \eta_1 \eta_2^2},
\]

where we used a Schwarz inequality, the norm bound \( \| G_2 G_2^* \| \leq \eta_2^{-4} \), and that \( \| G_1 \| \leq 1 \), \( \| (G_1 G_2 G_2) \| \leq \eta_1^{-1/2} \eta_2^{-3/2} \).

Note that the error term in the rhs. of (80) is smaller than our goal \( \rho^* (L \eta_1 \eta_2)^{-1} \) in (47), since for \( |z_1 - z_2| < \eta^* \) we have that

\[
\frac{1}{N \eta_2^2 \eta_1} \leq \frac{\rho^*}{L \eta_1 \eta_2}. \tag{82}
\]

This concludes the proof of the local law for \( (G^* G^*) \).

\[ \square \]

**Proof of the local law for** \( \langle G_1 G_2 AG_1 A' \rangle \) in (84). Consider the equation in (77) for \( G_3 = G_1 \), and \( B_1 = I, B_2 = A, B_3 = A' \), with \( \langle A \rangle = \langle A' \rangle = 0 \). Before proceeding with writing the equation for \( \langle G_1 G_2 AG_1 A' \rangle \), we bound two representative terms with a pre-factor \( N^{-1} \) in (77):

\[
\| (G_1^* G_1 G_2 AG_1 A') \| \leq \| G_1 G_1^* \|^{1/2} (G_2 G_2^* (G_1^*)^* A' G_1^* G_1^* G_2^* )^{1/2}
\]

where the first estimate we used [5, Lemma 5] to bound

\[
\| \langle G_2 G_1 A' G_1 G_1^* A \rangle \| \leq \langle G_1 G_1^* \rangle^{1/2} \langle G_2 G_2^* G_2^* G_2^* G_2^* \rangle^{1/2} \leq \frac{\sqrt{N \rho_1 \rho_2}}{\eta_1 \eta_2}, \tag{83}
\]

where we used that \( \langle G_2 G_1 A' \rangle = m_1 \langle G_2 AG_1 A' \rangle - m_1 \langle WG_1 G_2 AG_1 A' \rangle \)

+ \( m_1 \langle G_1 G_2 G_1 A' \rangle \)

+ \( m_1 \langle G_1 G_2 G_1 \rangle \langle G_1 A' \rangle + O_{\prec} \left( \frac{\rho_1 \rho_2}{\sqrt{L \eta_1 \eta_2}} \right) \tag{84}
\]

by [5, Lemma 5] again. The bound of all the other terms with a pre-factor \( N^{-1} \) is analogous and so omitted. Then, by (83) and (77), we conclude that

\[
\left[ 1 + O_{\prec} \left( \frac{1}{N \eta_2} \right) \right] \langle G_2 AG_1 A' \rangle = m_1 \langle G_2 AG_1 A' \rangle - m_1 \langle WG_1 G_2 AG_1 A' \rangle
\]

+ \( m_1 \langle G_1 G_2 G_1 A' \rangle \)

+ \( m_1 \langle G_1 G_2 G_1 \rangle \langle G_1 A' \rangle + O_{\prec} \left( \frac{\rho_1 \rho_2}{\sqrt{L \eta_1 \eta_2}} \right) \tag{84}
\]

where we used that

\[
\langle G_2 AG_1 A' \rangle = m_1 m_2 \langle AA' \rangle + O_{\prec} \left( \frac{\rho^*}{\sqrt{L}} \right), \quad \langle G_1 G_2 \rangle = \frac{m_1 m_2}{1 - m_1 m_2} + O_{\prec} \left( \frac{1}{N \eta_1 \eta_2} \right), \tag{85}
\]

and

\[
\| \langle G_1 A' \rangle \| \leq \frac{\rho_1}{N \sqrt{\eta_1}}. \tag{86}
\]

The local laws in (85) follow by (42), whilst the bound in (86) follows by (40).
Finally, using the bound $|\langle W G_1 G_2 A G_1 A \rangle| \prec \rho^* K^{-1/2} (\eta^*)^{-1}$ from [6, Theorem 5], we conclude that

$$\langle G_1 G_2 A G_1 A \rangle = \frac{m_1^2 m_2}{1 - m_1 m_2} (A A') + O_\prec \left( \frac{\rho^*}{\sqrt{L \eta_1 \eta_2}} \right),$$

where we recall that $K = N \eta_1 \rho^*$, and $L = N (\rho_1 \eta_1 \wedge \rho_2 \eta_2)$. This concludes the proof of the local law (48) for $\langle G_1 G_2 A G_1 A \rangle$. \hfill \Box

\textbf{Proof of the bound for $\langle G_1 G_2 G_2 A \rangle$ in (5).} Consider the equation in (77) for $G_3 = G_2$ and $B_1 = B_2 = I$, $B_3 = A$, with $\langle A \rangle = 0$. Proceeding similarly to (83) to estimate the error terms with a pre-factor $N^{-1}$, we conclude that

$$\langle G_1 G_2 G_2 A \rangle = m_1 \langle G_2 G_2 A \rangle - m_1 \langle W G_1 G_2 G_2 A \rangle + m_1 \langle G_1 - m_1 \rangle \langle G_1 G_2 G_2 A \rangle$$

$$+ m_1 \langle G_1 G_2 \rangle \langle G_2 G_2 A \rangle + m_1 \langle G_1 G_2 G_2 \rangle \langle G_2 A \rangle + O_\prec \left( \frac{\sqrt{\rho^*}}{L \eta_1 \eta_2} \right)$$

$$= O_\prec \left( \frac{\sqrt{\rho^*}}{L \eta_1 \eta_2} \right) \left( \frac{\rho^* \sqrt{\rho^*}}{K \sqrt{\eta_1 \eta_2}} \right),$$

where to go from the second to the third line we used that

$$|\langle W G_1 G_2 G_2 A \rangle| = O_\prec \left( \frac{\rho^* \sqrt{\rho^*}}{K \sqrt{\eta_1 \eta_2}} \right)$$

by (61), and that

$$|\langle G_2 G_2 A \rangle| = O_\prec \left( \frac{\sqrt{\rho^*}}{N \eta_2^{3/2}} \right), \quad |\langle G_1 G_2 \rangle| \prec \frac{\rho_1 \rho_2}{\eta_1 \eta_2}, \quad |\langle G_1 G_2 G_2 \rangle| \prec \frac{\sqrt{\rho_1 \rho_2}}{\sqrt{\eta_1 \eta_2}}. \quad (89)$$

The first bound in (89) follows by (45), whilst the second and the third bound follow by a simple Schwarz inequality. This concludes the proof of the bound for $\langle G_1 G_2 G_2 A \rangle$. \hfill \Box

The remaining statements of Proposition 3.4 involving transposed resolvents are proved similarly. For completeness we included those proofs in Appendix C. This concludes the proof of Proposition 3.4. \hfill \Box

4. CLT for resolvents

We now formulate the resolvent CLT identifying the joint distribution of $\langle (G - E G) \rangle$, $\langle (G - E G) A \rangle$ for multiple $z$’s and traceless $A$’s. An analogous result for only $(G - E G)$ factors was proven in [29]. Let $p \leq q \in \mathbb{N}$ and let $A_1, \ldots, A_p$ be matrices with $\langle A_i \rangle = 0$, $|A_i| \leq 1$ and let $a_i$ denote the vector of the diagonal elements of $A_i$. Let $z_1, \ldots, z_q \in \mathbb{C} \setminus \mathbb{R}$ be spectral parameters. We then set

$$G_i := G(z_i), \quad m_i := m(z_i), \quad \rho_i := \frac{1}{\pi} |3 m_i|, \quad \eta_i := |3 z_i|,$$

$$X_i := \langle [G_i - E G_i] A_i \rangle, \quad Y_i := \langle G_i - E G_i \rangle, \quad X_S := \prod_{i \in S} X_i, \quad Y_S := \prod_{i \in S} Y_i,$$

so that from the local laws in (40) we have the \textit{a priori} bounds.

$$|X_S| \prec \Psi_S, \quad |Y_S| \prec \Psi_S, \quad \Psi_S := \prod_{i \in S} \Psi_{i}, \quad \Psi_{i} := \frac{\rho_i}{N \eta_i^{1/2}} 1(i \leq p) + \frac{1}{N \eta_i} 1(i > p).$$

The following theorem identifies the leading terms of the joint moments of $X$’s and $Y$’s up to an error term that is smaller than the \textit{a priori} bounds.
Theorem 4.1. For any $\epsilon > 0$ and $1 \leq L := \min_i N \eta, \rho$, we have that
\[
\mathbb{E} X[p] Y[p] = \frac{1}{N^2} \sum_{P \in \text{Pair}((p, q))} \prod_{(i,j) \in P} V_{ij}^{\delta}(A_i, A_j) \prod_{(i,j) \in Q} V_{ij} + O\left(\frac{N^\delta \Psi}{\sqrt{L}}\right),
\]
(91)
where $\Psi := \Psi[q]$,
\[
V_{ij}^{\delta}(A_i, A_j) := \frac{m_i^2 m_j^2 \langle A_i, A_j \rangle}{1 - \sigma_i \sigma_j m_i^2} \left[ \kappa_4 m_i^2 m_j^2 + \tilde{w}_2 m_i^2 m_j^2 \right] \langle a_i, a_j \rangle
\]
\[
V_{ij} := \frac{m_i' m_j'}{(1 - \sigma_i m_i)^2} + \frac{\sigma_i m_i' m_j'}{(1 - \sigma_i m_i)^2} + \frac{\kappa_4}{2} (m_i^2) (m_j^2) + \tilde{w}_2 m_i m_j
\]
(92)
and $\text{Pair}(S)$ denotes the set of pairings of a base set $S$. Moreover, the expectation $\mathbb{E} G$ is given by
\[
\langle \mathbb{E} G_i \rangle = m_i + \frac{1}{N} \left( \frac{m_i^2}{1 - \sigma_i m_i^2} + \tilde{w}_2 m_i m_i + \frac{\kappa_4}{2} (m_i^2) \right) + O\left(\frac{N^\delta \Psi_i}{L^{1/2}}\right).
\]
(93)

Note that the first (leading) term in (91) has a natural size of order $\Psi$ whenever $(3z_i)(3z_j) < 0$ for every $\{i, j\}$ in the pairings $P, Q$.

Within the proof of Theorem 4.1 we use the classical cumulant expansion in the form
\[
\mathbb{E} w_{ab} f(W) = \sum_{k=1}^{R-1} \sum_{\alpha \in \{ab, ba\}^k} \frac{\kappa(ab, \alpha)}{k!} \mathbb{E} \partial_{\alpha} f(W) + \Omega_R,
\]
(94)
where $\kappa(ab, \alpha)$ denotes the joint cumulant of $w_{ab}, w_{\alpha_1}, \ldots, w_{\alpha_k}$ for $\alpha = (\alpha_1, \ldots, \alpha_k)$. Here for any cut-off index $R$ the error term $\Omega_R$ has an explicit integral representation [22, Proposition 3.2]. For our application, where $f(W)$ is a product of resolvents, the error term can easily be estimated by $|\Omega_R| \lesssim N^{-(R+1)/2}$. This is due to the fact that the $k$-th cumulant scales like $N^{-k/2}$, and each derivative creates an additional resolvent entry which can be estimated by $O(1)$ due to the single resolvent local law. In the sequel we will omit the cutoff $R$ from the formulas and we simply write a cumulant expansion with a formally infinite sum over $k$, but technically we always estimate the truncated sum.

Proof of Theorem 4.1. Recalling from (63) that
\[
G = m - mWG + m(G - m)G + m m G^2 G N + m \tilde{w}_2 \sqrt{N} \text{diag}(G)G,
\]
(95)
it follows that (with $\eta = |3z|, \rho = \rho(z)$)
\[
(1 - m^2) (G - m) = -m(WG) + m(G - m)^2 + \frac{\sigma m}{N} (G^4 G) + m \tilde{w}_2 \sqrt{N} \langle \text{diag}(G)G \rangle
\]
\[
= -m(WG) + \frac{m}{N} \left( \frac{\sigma m^2}{1 - \sigma m^2} + \tilde{w}_2 m^2 \right) + O_{\epsilon'} \left( \frac{\rho}{N \eta L^{1/2}} \right)
\]
(96)
from (43) and $\langle \text{diag}(G)G \rangle = m^2 + O_{\epsilon'} (\sqrt{\rho / N \eta})$ due to (4). Using a cumulant expansion we prove below that
\[
\mathbb{E} \langle WG \rangle = \frac{1}{N} \sum_{k \geq 2} \sum_{ab} \sum_{\alpha \in \{ab, ba\}^k} \frac{\kappa(ab, \alpha)}{k!} \mathbb{E} \partial_{\alpha} G_{ba} = -\kappa_4 m^4 N + O\left(\frac{\rho}{N^2 \eta} + \frac{\rho^{3/2}}{N^{3/2} \eta^{1/2}}\right),
\]
(97)
where we ignored the irrelevant error term $\Omega_R$. Note that the $k = 1$ summand has been cancelled by the variance term which is included in the definition of the “underline” renormalisation. Then, the first claim in (93) follows immediately from (97) together with (96).
Similarly, from (95) we obtain

\[
(1 - m(G - m))(GA) = -m(WGGA) + \frac{\sigma m}{N} (G^2 GA) + m \frac{\sqrt{\rho}}{N} (\text{diag}(G) GA)
\]

\[
= -m(WGGA) + O\left(\frac{\rho^{1/2}}{N^3/2}\right)
\]  

(98)

from (48) and

\[
\langle \text{diag}(G) GA \rangle = m \langle GA \rangle + \langle \text{diag}(G - m) GA \rangle = O\left(\sqrt{\frac{\rho}{N^3/2}}\right).
\]

For the underlined term in (98) it follows exactly as in (97) that

\[
E(WGGA) = \frac{1}{N} \sum_{k \geq 2} \sum_{ab} \sum_{\alpha \in \{a, b, a^c\}} \frac{\kappa(ab, \alpha)}{k!} E \partial_{\alpha}(GA)_{ba} = O\left(\frac{\rho}{N^3/2} + \frac{\rho^{3/2}}{N^{3/2} \eta^{1/2}}\right),
\]

(99)

where the term corresponding to \( \kappa_4 \) vanishes due to \( \langle A \rangle = 0 \). Together with (98), the second claim in (93) is also proven. This concludes the computation of the expectation.

We will now prove an asymptotic Wick theorem and explicitly compute the variance. Using (40) and (98)–(99) we replace \( X_1 = \langle [G_1 - E] G_1 \rangle A_1 \) with its leading term \( WG_1 A_1 \), i.e.

\[
\langle [G_1 - E] G_1 \rangle A_1 = -m(WG_1 A_1) + m E(WG_1 A_1) + O\left(\frac{\rho}{N^3/2}\right)
\]

\[
= -m(WG_1 A_1) + O\left(\frac{\rho}{N^3/2}\right).
\]

Then for

\[
E X_{[p]} Y_{[p]} = -m_1 E(WG_1 A_1) X_{[1,p]} Y_{[p]} + O\left(\frac{N^2 \psi}{L^{1/2}}\right)
\]

(100)

we perform a cumulant expansion of \( (WG_1 A_1) \) to obtain

\[
E X_{[p]} Y_{[p]} = \sum_{\kappa \in \{1\}} m_1 E(WG_1 A_1) G_{[p]} W G_{[p]} A_{[1]} Y_{[p]} + O\left(\frac{N^2 \psi}{L^{1/2}}\right)
\]

\[
+ \sum_{\kappa \in \{p\}} m_1 E(WG_1 A_1) G_{[p]} W G_{[p]} X_{[p]} Y_{[p]}
\]

\[
- \sum_{k \geq 2} \sum_{ab} \sum_{\alpha \in \{a, b, a^c\}} \frac{\kappa(ab, \alpha)}{k! N} E \partial_{\alpha} \left[ m_1 (G_1 A_1)_{ba} X_{[p]} Y_{[p]} \right]
\]

(102)

We note that for arbitrary matrices \( U, V \) independent of \( W \) we have

\[
E(WU)(WV) = \frac{1}{N^2} \left( \langle UV \rangle + \sigma \langle UV^t \rangle + \widetilde{w}_2 \langle \text{diag}(U) \text{diag}(V) \rangle \right),
\]

(103)

so that it follows that

\[
N^2 E(WG_1 A_1)(G_i W G_1 A_i)
\]

\[
= \langle G_1 A_1 G_i A_i \rangle + \sigma \langle G_1 A_1 G_i^t A_i^t \rangle + \widetilde{w}_2 \langle \text{diag}(G_1 A_1) \text{diag}(G_i A_i) \rangle
\]

\[
= \frac{m_1 m_2^2}{1 - m_1 m_1} \frac{\langle A_1 A_i \rangle}{1 - \sigma m_2 m_1} + \frac{\sigma m_1 m_2^2}{1 - \sigma m_1 m_1} \frac{\langle A_1 A_i \rangle}{\sqrt{\rho}} + \frac{m_1 m_2^2}{\sqrt{\rho}} \langle A_1 a_i \rangle + O\left(\frac{N^2 \psi}{L^{1/2}}\right)
\]

(104)

from (48), (50), and

\[
\langle \text{diag}(G_1 A_1) \text{diag}(G_i A_i) \rangle = m_1 \langle \text{diag}(a_1) G_i A_i \rangle + O\left(\frac{1}{\sqrt{L}} \frac{\sqrt{\rho_1}}{\sqrt{\rho_2}}\right)
\]

\[
= m_1 m_2^2 \langle a_1 a_i \rangle + O\left(\frac{1}{\sqrt{L}} \frac{\sqrt{\rho_1}}{\sqrt{\rho_2}}\right)
\]
due to the second bound in (46) and the second local law in (42). Similarly, for the second line on the rhs. of (102) we obtain
\[ N^2 \mathbb{E} (\tilde{W} G_1 A_1 | G_i, \tilde{W} G_i) = \langle G_1 A_1 G_i^2 \rangle + \sigma \langle G_1 A_1 (G_i^2)^\dagger \rangle + \tilde{w}_2 \langle \text{diag}(G_1 A_1) \text{diag}(G_i^2) \rangle = \mathcal{O}_2 \left( \frac{N^2 \Psi(1+i)}{\sqrt{L}} \right). \]
(105)

from (51) and
\[ \langle \text{diag}(G_1 A_1) \text{diag}(G_i^2) \rangle = m_1 \langle \text{diag}(\alpha_i) G_i^2 \rangle + \mathcal{O}_\prec \left( \frac{1}{\sqrt{L \eta_i}} \right) = \mathcal{O}_\prec \left( \frac{1}{\sqrt{L \eta_i}} \right) \]
due to (45) and the first local law in (46).

It remains to consider the third line in (102) where due to the Leibniz rule many terms can arise from the derivative. For \( k \geq 2 \) the derivative may act on \( G_1 \) or any of the \( X_i, Y_i \) and we consider the corresponding terms separately as in
\[ \sum_{k \geq 2} \sum_{a b} \sum_{\alpha \in \{a b, b a\}^k} \frac{\kappa(ab, \alpha)}{k!N} \partial_\alpha \left[ m_1 \langle G_1 A_1 \rangle_{ba} X_{p \setminus \{1\}} Y_{(p, q)} \right] \]
\[ = \sum_{k \geq 2} \sum_{|P_X \cup P_Y| \leq k} \sum_{X(1,p) \setminus S(P_X) \subset (p,q)} X_{(1,p) \setminus S(P_X) \setminus S(P_Y)} \Xi_k(P_X, P_Y), \]
\[ \Xi_k(P_X, P_Y) := \sum_{a b} \sum_{\alpha} \frac{\kappa(ab, \alpha)}{k!N} \prod_{i \in S(P_X)} \partial_{\alpha_i} \left[ m_1 \langle G_1 A_1 \rangle_{ba} \right] \prod_{i \in S(P_X)} \partial_{\alpha_i} X_i \prod_{i \in S(P_Y)} \partial_{\alpha_i} Y_i \frac{1}{k_i!} \cdot \]
(106)

where \( P_X, P_Y \) are unordered multisets with support \( S(P_X) \subset (1, p) \), \( S(P_Y) \subset (p, q) \). The last summation \( \sum_{\alpha} \) indicates the summation over tuples \( \alpha_i \in \{a b, b a\}^{k_i}, \alpha_i \in \{a b, b a\}^{k_i} \) with \( k_i \geq 1 \) denoting the multiplicity of \( i \) in \( S(P_X \cup P_Y) \) and \( k_1 := k - |P_X \cup P_Y| \geq 0 \). We will prove below that
\[ \Xi_k(P_X, P_Y) = \mathcal{O}_\prec \left( \frac{\Psi(1) \cup S(P_X \cup P_Y)}{\sqrt{L}} \right) \]
\[ - \frac{m_1 m_3}{N^2} \frac{\kappa_4}{2} (k, P_X, P_Y) = (3, \{i, i, \emptyset\}, i \in (1, p)) \]
(107)

By combining (102), (104), (105), (106) and (107) we obtain from induction on the number of \( X \)-factors
\[ \mathbb{E} X_{p \setminus \{1\}} Y_{(p, q)} = \mathbb{E} Y_{(p, q)} \frac{1}{N^p} \mathbb{E} \prod_{P \in \text{Pair}(p)} \prod_{(i, j) \in P} V_{\psi(i, j)}(A_i, A_j) + \mathcal{O} \left( \frac{N^4 \Psi}{L^{1/2}} \right). \]
(108)

Here we used that for \( X_{1,p) \setminus S(P_X) \) and \( Y_{(p,q) \setminus S(P_Y)} \) we have the high probability a priori bounds \( \left| X_{1,p) \setminus S(P_X) \right| < \Psi(1,p) \setminus S(P_X) \) and \( \left| Y_{(p,q) \setminus S(P_Y)} \right| < \Psi(p,q) \setminus S(P_Y) \) and therefore
\[ \mathbb{E} \left| X_{(1,p) \setminus S(P_X)} Y_{(p,q) \setminus S(P_Y)} \right| \mathcal{O}_\prec \left( \frac{\Psi(1) \cup S(P_X \cup P_Y)}{\sqrt{L}} \right) \leq \mathbb{E} \mathcal{O}_\prec \left( \frac{\Psi(p,q)}{\sqrt{L}} \right) \leq \frac{N^4 \Psi}{L^{1/2}}. \]

In order to complete the proof of the theorem it remains to compute \( \mathbb{E} Y_{(p,q)} \). For convenience of notation we relabel \( Y_{(p,q)} \) to \( Y_{(r)} \) with \( r = q - p + 1 \) and obtain, analogously to (102),
\[ \mathbb{E} Y_{(r)} = \sum_{i \in \{2, r\}} \frac{m_1^i}{m_1} \mathbb{E} \mathbb{E} (\tilde{W} G_1 | G_i, \tilde{W} G_i) Y_{(r) \setminus \{1, i\}} = \frac{2 \kappa_4}{N} m_1^3 m_1^3 Y_{[2, r]} \]
\[ - \sum_{k \geq 2} \sum_{a b} \sum_{\alpha \in \{a b, b a\}^k} \frac{\kappa(ab, \alpha)}{k!N} \mathbb{E} \partial_\alpha \left[ \frac{m_1^i}{m_1} \langle G_1 \rangle_{ba} Y_{(1, r)} \right]. \]
(109)
For the first term on the rhs. of (109) we obtain with (103) that
\[
N^2 \mathbb{E} \langle \widetilde{WG}_1 \rangle \langle G_i \widetilde{WG}_i \rangle \\
= \langle G_i^2 \rangle + \sigma \langle G_i \rangle^2 + \tilde{w}_2 \langle \text{diag}(G_i) \rangle \text{diag}(G_i^2) \\
= \frac{m_1 m_i'}{(1 - m_1 m_i)^2} + \frac{\sigma m_1 m_i'}{(1 - \sigma m_1 m_i)^2} + \tilde{w}_2 m_1 m_i' + O_\prec \left( \frac{\rho_1}{\sqrt{L \eta_1 \eta_i}} \right),
\]
where in the last step we used (47), (49), and
\[
\langle \text{diag}(G_i) \rangle \text{diag}(G_i^2) = m_1 \langle G_i^2 \rangle + O_\prec \left( \frac{\rho_1}{L^{1/2} \eta_i} \right) = m_1 m_i' + O_\prec \left( \frac{\rho_1}{L^{1/2} \eta_i} \right)
\]
due to (41), the first local law in (42), and the first local law in (46).

For the second line of (109) we distribute the derivative according to the Leibniz rule as
\[
\sum_{ab} \sum_{\alpha \in \{ab, ba\}^k} \frac{\kappa \langle ab, \alpha \rangle}{k!} \mathbb{E} \partial_{\alpha} \left[ \frac{m_i}{m_1} \langle G_i \rangle_{ba} Y_{(1, r)} \right] = \sum_{|P_Y| \leq k} \mathbb{E} Y_{(1, r) \setminus S(P_Y)} \Phi_k(P_Y),
\]
\[
\Phi_k(P_Y) := \sum_{ab} \sum_{\alpha} \frac{\kappa \langle ab, \alpha \rangle}{N} \left[ \frac{m_i}{m_1} \partial_{\alpha} \langle G_i \rangle_{ba} \right] \left( \prod_{i \in S(P_Y)} \frac{\partial_{y_i} Y_i}{k_i!} \right),
\]
where $P_Y$ is a multiset with support $S(P_Y) \subset \{1, r\}$, and the summation $\sum_{\alpha}$ indicates the summation over tuples $\alpha \in \{ab, ba\}^{k_1}$, $\alpha_i \in \{ab, ba\}^{k_i}$ with $k_i \geq 1$ denoting the multiplicity of $i \in S(P_Y)$ and $k := k - |P_Y| \geq 0$. Similarly to (109) (but in high probability sense) we prove below that
\[
\Phi_k(P_Y) = O_\prec \left( \frac{\Psi_{1(1) \setminus S(P_Y)}}{\sqrt{L}} \right) - \langle m_i \rangle ' (\langle m_i \rangle ')^t \frac{\kappa_4}{2N^2} \mathbb{1} \left( (k, P_Y) = (3, \{i, i\}), i \in (1, r) \right)
\]
\[
- \frac{\kappa_4}{N} m_1 m_i' \mathbb{1} \left( (k, P_Y) = (3, \emptyset) \right).
\]
By combining (109), (110), (111) and (112) we conclude
\[
\mathbb{E} Y_{(r)} = \frac{1}{N^r} \sum_{Q \in \Psi_{\text{pair}(r)}} \prod_{(i, j) \in Q} V_{ij} + O \left( \frac{N^r \Psi_{(r)}}{L^{1/2}} \right).
\]

Therefore the claim (91) follows immediately from combining (108) and (113) and the proof of the Theorem 4.1 is complete, modulo the proofs of (107) and (112) which we present after Lemma 4.2 below.

\[\square\]

4.1. Auxiliary calculations: Proof of (97), (107) and (112).

Proof of (97). For $k = 2$ the summation in (97) has either two or none diagonal $G$’s and from (41) we estimate the corresponding terms by
\[
N^{-5/2} \sum_{ab} |G_{ba}|^3 < N^{-3/2} + \frac{\rho^{3/2}}{N^{2\eta_1/2}}
\]
and
\[
N^{-5/2} \sum_{ab} G_{ba} G_{ab} |G_{ba}| < N^{-3/2} + \frac{\rho^{3/2}}{N^{2\eta_1/2}}.
\]
Here the first estimate uses only $|G_{ab}| \ll 1 (a = b) + \sqrt{\rho/N \eta}$ while the second one uses $G = m + (G - m)$ and the isotropic resummation procedure. More precisely, by this we mean the idea of
summing up the free indices into constant vectors, i.e.
\[
\sum_{ab} G_{ab} \delta_{aa} G_{ba} = m^2 G_{11} + m \sum_a G_{1a} \mathcal{O}\left(\sqrt{\frac{\rho}{N\eta}}\right) + m \sum_b G_{b1} \mathcal{O}\left(\sqrt{\frac{\rho}{N\eta}}\right) + \sum_{ab} \mathcal{O}\left(\frac{\rho}{N\eta}\right) G_{ba}
\]
(114)

where we used a Schwarz inequality, and in the last step the isotropic local law for the all-one vector \( \mathbf{1} = (1, \ldots, 1) \) of norm \( \|\mathbf{1}\| = \sqrt{N} \). Thus we obtain a bound \( N^{-3/2} + N^{-2} \rho^{3/2} / N^{3/2} / \eta^{3/2} \) for the \( k = 2 \) terms in (97).

Next, we consider the \( k = 3 \) terms which give a contribution of \( \rho N^{-2} \eta^{-1} \) whenever there are at least two off-diagonal \( G \)'s. In order to achieve only diagonal \( G \)'s, \( \alpha \) is necessarily one of \( (ab, ab, ab) \), \( (ba, ab, ba) \), or \( (ba, ba, ab) \), for which we obtain \( \kappa(ab, ba, ba, ab) = \kappa_4 / N^3 \) for \( a \neq b \). The derivative then is given by
\[
\partial_{ab} G_{ba} = -\partial_{ab} \partial_{bd} G_{dab} = 2\partial_{ba} G_{ba} G_{ab} G_{ab} = -2G_{ab}^2 G_{ba}^2 + \cdots
\]
(115)

where the neglected terms contain two off-diagonal \( G \)'s and can hence be neglected. Therefore the \( k = 3 \) contribution of (97) is given by
\[
-2 \frac{3}{3!} \frac{1}{N^3} \sum_{ab} G_{ab}^2 G_{ba}^2 = -\kappa_4 \frac{N^3}{N^3} m^4 + \mathcal{O}\left(\frac{\rho}{N^2 \eta} + \frac{\rho^{3/2}}{N^{3/2} \eta^{3/2}}\right).
\]

By estimating the \( k \geq 4 \) contribution trivially via \( |G_{ab}| < 1 \) this concludes the proof of (97). \( \square \)

**Lemma 4.2** (Auxiliary a priori estimates). For \( X_i, Y_i \) from (90) and their derivatives \( \partial_{\alpha} X_i, \partial_{\alpha} Y_i \) for any multi-index \( \alpha \) we have the high probability a priori estimates
\[
|\partial_{\alpha} X_i| < \frac{\rho^3/2}{N^{1/2} \eta_i}, \quad |\partial_{\alpha} Y_i| < \frac{1}{N^{1/2} \eta_i},\]
(116)

and the more precise expansions for the first and second order derivatives
\[
\partial_{\alpha} Y_i = -m_i^2 \frac{\delta_{aa}}{N} + \mathcal{O}\left(\frac{\rho_i^3/2}{N \eta_i^{3/2}}\right), \quad \partial_{\alpha} \partial_{\alpha} Y_i = 2m_i m_i' \frac{\delta_{aa} \delta_{bb}}{N} + \mathcal{O}\left(\frac{\rho_i^3/2}{N \eta_i^{3/2}}\right).
\]
(117)

Moreover, we have the expansions
\[
\partial_{\alpha} X_i = -m_i^2 \frac{(A_i)_{\alpha a}}{N} + \mathcal{O}\left(\frac{\rho_i}{N^{3/2} \eta_i}\right), \quad \partial_{\alpha} \partial_{\alpha} X_i = m_i^2 \frac{(A_i)_{\alpha a} \delta_{aa} + (A_i)_{\beta c} \delta_{ba}}{N} + \mathcal{O}\left(\frac{\rho_i}{N^{3/2} \eta_i}\right),
\]
(118)
in variance sense.

**Proof.** We first establish an isotropic local law in variance sense using (59) of the form
\[
(GAG)_{xy} = m^2 A_{xy} + \mathcal{O}\left(\frac{\rho}{N^{1/2} \eta}\right)
\]
(119)
which is proved analogously to (73). The claims (117)–(118) then follow directly from (119), the first local law in (46), and
\[
\partial_{ab} (GB) = -\frac{(GB)_{ba}}{N}
\]
(120)
and
\[
\partial_{ab} \partial_{cd} (GB) = -\frac{(GB)_{ba}}{N} \equiv G_{ba} (GB)_{da} + G_{ba} (GB)_{da}.
\]
(121)
The claim (16) follows inductively by the second local law in (46) since each additional derivative just adds an additional factor of $G$ which is at most of order 1. □

**Proof of (107).** We prove (107) by considering the following five cases which cover all possibilities: (i) $k_1$ odd, $k \geq 4$, (ii) $k_1$ even, $k \geq 3$, (iii) $k = 3$, $k_1 = 1$ (iv) $k = 3$, $k_1 = 3$ and (v) $k = 2$. Before considering each case separately, we outline a few ideas that are used repeatedly in the argument. The first idea is that we often replace diagonal resolvents $G_{aa}$ and $(GA)_{aa}$ using the isotropic local law $\langle x, G y \rangle = m \langle x, y \rangle + O_{\infty} \left( \frac{\rho}{N \eta} \right)$ in order to make the leading term independent of the summation index $a$. For example, for $\sum_a G_{aa} G_{ax}$ this allows us to sum up the index $a$ into the constant vector $1 = (1, \ldots, 1)$ of norm $\sqrt{N}$, effectively gaining a factor of $\sqrt{\rho/N \eta}$ over the naive estimate since

$$\sum_a G_{aa} G_{ax} = m G_{1x} + O_{\infty} \left( \frac{\rho}{N \eta} \sum_a |G_{ax}| \right) = O_{\infty} \left( N \frac{\rho}{N \eta} \right).$$

The second idea is that for off-diagonal resolvents we use a Schwarz inequality, followed by the Ward identity to effectively also gain a factor of $\sqrt{\rho/N \eta}$ over the naive estimate, e.g.

$$\left| \sum_a |G_{ax}| \right| \leq \sqrt{N} \sum_a |G_{ax}|^2 = \sqrt{N} \sqrt{|(G^2 G)_{xx}|} = \sqrt{\frac{N}{\eta}} \sqrt{\frac{\rho}{N \eta}}.$$

Finally, we also frequently use a simple parity consideration to count off-diagonal resolvents since the local law gives a stronger estimate for them. For an odd number of $G$'s, each evaluated in one of the entries $aa, bb, ab, ba$ with $a \neq b$ in total occurring equally often, at least one of the $G$'s has to be off-diagonal.

**Case (i), $k_1$ odd, $k \geq 4$.** In this case we estimate $|\partial x_1 (G_1 A_1)_{ba}| \prec 1$ in the definition of $\Xi$ in (106) and obtain from (16) that

$$|\Xi_k (P_X, P_Y)| \prec N^{-(k+3)/2} N^2 \Psi_S (P_X \cup P_Y) \lesssim N^{-(k-3)/2} \Psi_{(1) \cup S} (P_X \cup P_Y),$$

from $N^{-1} \lesssim \rho_1^{1/2} N^{-1} \eta_1^{-1/2}$, confirming (107).

**Case (ii), $k_1$ even, $k \geq 3$.** Since $k_1 = |x_1|$ is odd it follows by parity that at least one $G$ or $GA$ factor is off-diagonal, hence by the local law we have that

$$|\partial x_1 (G_1 A_1)_{ba}| \prec |(G_1)_{ab}| + |(G_1 A_1)_{ba}|$$

and therefore, by (16) and a Ward-estimate it follows that

$$|\Xi_k (P_X, P_Y)| \prec N^{-(k+3)/2} \sum_{ab} |(G_1 A_1)_{ba}| + |(G_1)_{ab}| \Psi_S (P_X \cup P_Y)$$

$$\lesssim N^{-(k-3)/2} \sqrt{\frac{\rho_1}{N \eta_1}} \sqrt{\frac{\rho}{N \eta}} \Psi_S (P_X \cup P_Y) \lesssim N^{-(k-2)/2} \Psi_{(1) \cup S} (P_X \cup P_Y),$$

confirming (107).

**Case (iii), $k = 3, k_1 = 3$.** The three derivatives acting on $(G_1 A_1)_{ab}$ results in one $G_1 A$ and three $G_1$ factors with a total of four $a$ and four $b$ indices. By using the local law we replace each $G_1$ by $m_1$ and obtain

$$|\Xi_k (0, 0)| \lesssim N^{-3} \sum_{ab} m_1^3 (A_1)_{aa} + N^{-3} \sum_{ab} m_1^3 (A_1)_{ab} + O_{\infty} \left( \frac{\Psi_{(1)}}{\sqrt{L}} \right)$$

$$\lesssim N^{-2} \sqrt{\sum_{ab} |(A_1)_{ab}|^2} + \frac{\Psi_{(1)}}{\sqrt{L}} = N^{-3/2} \sqrt{(A_1 A_1^*)} + \frac{\Psi_{(1)}}{\sqrt{L}},$$

again confirming (107).
Case (iv), $k = 3$, $k_1 = 1$. For $k_1 = 1$ the derivative of $(G_1 A_1)_{ba}$ is given by
\[
\sum_{\alpha_1} \partial_{\alpha_1} (G_1 A_1)_{ba} = -(G_1)_{ba} (G_1 A_1)_{ba} - (G_1)_{bb} (G_1 A_1)_{aa} = -m_1^2 (1 + \delta_{ba}) (A_1)_{aa} + O_{\infty} \left( \sqrt{\rho_1 N \eta_i} \right).
\]
If $P_X = \{i, j\}$ for some $i \in (1, p]$, then we obtain from (118) that
\[
\partial_{ab, ba} X_i = \partial_{ba, ab} X_i = m_1^3 (A_1)_{bb} + (A_1)_{aa} + O_{\infty} \left( \rho_1 N^{3/2} \eta_i \right),
\]
while both the $\partial_{ab, ba}$ and $\partial_{ba, ab}$ derivatives lead to delta functions $\delta_{ab}$ and are therefore lower order after summation, and thus
\[
\Xi_3(\{i, j\}, \emptyset) = -m_1^3 m_i^3 \sum_{ab} \kappa(ab, ba, ab, ba) \frac{(A_1)_{aa}}{N} (A_1)_{ba} + (A_1)_{bb} + O_{\infty} \left( \frac{\Psi_{\{i,j\}}}{L^{1/2}} \right)
\]
\[
= -\frac{K_4}{N^2} m_1^3 m_i^3 \left( \langle a_1 a_i \rangle + \langle A_1 \rangle \langle A_i \rangle \right) + O_{\infty} \left( \frac{\Psi_{\{1, i\}}}{L^{3/2}} \right)
\]
\[
= -\frac{K_4}{N^2} m_1^3 m_i^3 \langle a_1 a_i \rangle + O_{\infty} \left( \frac{\Psi_{\{1, i\}}}{L^{3/2}} \right),
\]
giving the leading contribution to (107).

If $P_Y = \{i, j\}$ for some $i \in (p, q]$, then we obtain from the leading term of (117) that
\[
\partial_{ab, ba} Y_i = \partial_{ba, ab} Y_i = m_1 m_i^2 \frac{2}{N} + O_{\infty} \left( \frac{\rho_1^{1/2}}{(N \eta_i)^{3/2}} \right),
\]
with the other derivatives again being lower order, hence
\[
\Xi_3(\emptyset, \{i, j\}) = -m_1^3 m_i^3 \sum_{ab} \kappa(ab, ba, ab, ba) \frac{(A_1)_{aa}}{N} \frac{2}{N} + O_{\infty} \left( \frac{\Psi_{\{1, i\}}}{L^{1/2}} \right)
\]
\[
= -\frac{2 K_4}{N^2} m_1^3 m_i^3 \langle A_1 \rangle + O_{\infty} \left( \frac{\Psi_{\{1, i\}}}{L^{1/2}} \right) = O_{\infty} \left( \frac{\Psi_{\{1, i\}}}{L^{3/2}} \right).
\]

Finally, if $P_X \cup P_Y = \{i, j\}$ for some $1 < i < j$, then we either obtain a $\delta_{ab}$ from (117) or a $(A_1)_{ba}$ from (118) and therefore due to
\[
\sum_{ab} |(A_1)_{ba}| \leq N \sqrt{\sum_{ab} |(A_1)_{ba}|^2} = N^{3/2} \sqrt{\langle A_1 A_1^* \rangle}
\]
the leading term is at most of size $\Psi_{\{1, i, j\}} L^{-1/2}$ and we obtain
\[
|\Xi_3(\{i, j\}, \emptyset)| + |\Xi_3(\emptyset, \{i, j\})| + |\Xi_3(\{i\}, \{j\})| = O_{\infty} \left( \frac{\Psi_{\{1, i, j\}}}{L^{1/2}} \right).
\]
Here we used (39a), so that in case $P_X = \{i, j\}$ with $i \neq j$, the error terms from (118) can be multiplied. This concludes the proof of (107) for the case $k_1 = 1$, $k = 3$.

Case (v), $k = 2$. In case $k = 2$ there are five sub-cases to consider; $k_1 = 2$, $P_X = \{i, i\}$, $P_Y = \{i, i\}$, $k_1 = 1$ or $|S(P_X \cup P_Y)| = 2$.

If $k_1 = 2$, then the derivative is given by
\[
\partial^2 (G_1 A_1)_{ba} = (G_1)(G_1)(G_1 A_1)
\]
with three $a$ and three $b$ indices, so that by parity either all three matrices have indices $ab, ba$, or only one with the remaining two having $aa, bb$. For all three matrices having $ab, ba$ indices, we can gain two factors of $\sqrt{\rho_1 / N \eta_i}$ via Ward-estimates over the naive size $N^{3/2}$ in order to obtain $\rho_1 N^{-3/2} \eta_i^{-1} \leq$
\( \Psi_{11} L^{-1/2} \). If two matrices have indices \( aa, bb \), then we replace one diagonal resolvent by \( m \) and estimate, for example,

\[
N^{-5/2} \sum_{ab} (G_1)_{bb} (G_1 A_1)_{aa} (G_1)_{ba} = m_1 N^{-5/2} \sum_a (G_1 A_1)_{aa} (G_1)_{1a} + O_{\prec} \left( \frac{\rho_1}{N^{3/2} \eta_1} \right)
\]

\[
= O_{\prec} \left( \frac{\rho_1^{1/2}}{N^{3/2} \eta_1^{1/2}} \right) = O_{\prec} \left( \frac{\Psi_{11}}{L^{1/2}} \right),
\]

and similarly for all other index distributions. Thus we obtain

\[
|\Xi_2(\emptyset, 0)| \lesssim \frac{\Psi_{11}}{L^{1/2}}.
\]

Next, if \( P_X = \{i, i\} \), then we obtain from (18) that

\[
\Xi_2(i, i, \emptyset) = \sum_{ab} N \frac{m_1 (G_1 A_1)_{ba}}{N} m_1 \frac{(A_i)_{aa} + (A_i)_{bb}}{N} + O_{\prec} \left( \frac{\Psi_{1i, 1i}}{L^{1/2}} \right)
\]

\[
= m_1 m_1 \frac{K^2}{N^{3/2}} \left( \sum_a (G_1 A_1)_{ba} + \sum_a (G_1 A_1)_{aa} \right) + O_{\prec} \left( \frac{\Psi_{1i, 1i}}{L^{1/2}} \right) = O_{\prec} \left( \frac{\Psi_{1i, 1i}}{L^{1/2}} \right)
\]

using that

\[
\|a_i\| = \|\text{diag} A_i\| \leq N^{1/2} \sqrt{\langle A_i^* A_i \rangle}.
\]

The case \( P_Y = \{i, i\} \) is completely analogous, except that using (17) the constant 1 vector is summed up instead of \( a_i \), and we obtain

\[
|\Xi_2(\emptyset, \{i, i\})| \lesssim \frac{\Psi_{1i, 1i}}{L^{1/2}}.
\]

The case \( k_1 = 1 \) can be estimated by

\[
|\Xi_2(\{i\}, \emptyset)| \lesssim N^{-7/2} \sum_{ab} \left| (G_1)_{ba} (G_1 A_1)_{aa} + (G_1)_{ba} (G_1 A_1)_{ba} \right| (G_1 A_i G_i)_{ba} \]

\[
\lesssim N^{-7/2} \sum_a \left| (G_1 A_i G_i)_{1a} \right| + N^{-7/2} \sum_{ab} \sqrt{\frac{\rho_1}{N \eta_1}} \left| (G_1 A_i G_i)_{ba} \right|
\]

\[
\lesssim \frac{\rho_1^{1/2}}{N \eta_1^{1/2}} \frac{\rho_i}{N^{3/2} \eta_i} \lesssim \frac{\Psi_{1i, 1i}}{L^{1/2}}
\]

and similarly

\[
|\Xi_2(\emptyset, \{i\})| \lesssim N^{-7/2} \sum_{ab} \left| (G_1)_{ba} (G_1 A_1)_{aa} + (G_1)_{ba} (G_1 A_1)_{ba} \right| (G_i^2)_{ba} \]

\[
\lesssim N^{-7/2} \sum_a \left| (G_i^2)_{1a} \right| + N^{-7/2} \sum_{ab} \sqrt{\frac{\rho_1}{N \eta_1}} \left| (G_i^2)_{ba} \right| \lesssim \frac{\rho_1^{1/2}}{N \eta_1^{1/2}} \frac{\rho_i^3}{N^{3/2} \eta_i^{3/2}} \lesssim \frac{\Psi_{1i, 1i}}{L^{1/2}}
\]

from (20).
For the final case \(|S(P_X \cup P_Y)| = 2\) we estimate for \(i \neq j\)
\[
|\Xi_2(\{i, j\}, \emptyset)| \lesssim \frac{1}{N^{3/2}} \sum_{ab} |(G_1 A_1)_{ba}| \left( \frac{|(A_1)_{ba}|}{N} + O^2 \left( \frac{\rho_{ij}}{N^{3/2} |\eta_j|} \right) \right)
\times \left( \frac{|(A_1)_{ba}|}{N} + O^2 \left( \frac{\rho_{ij}}{N^{3/2} |\eta_j|} \right) \right)
\]
\[
= O^2 \left( N^{-1/2} \sqrt{\frac{\rho_1}{N |\eta_i|}} \left( \frac{1}{N^{3/2} |\eta_j|} + \frac{\rho_{ij}}{N |\eta_i|^{1/2}} \right) \right) = O^2 \left( \Psi_{(1, i, j)} \right)
\]
\[
|\Xi_2(\emptyset, \{i, j\})| \lesssim N^{-5/2} \sum_{ab} |(G_1 A_1)_{ba}| \left( \frac{\delta_{ba}}{N} + O(\frac{\rho_{ij}}{(N |\eta_i|)^{3/2}}) \right)
\times \left( \frac{\delta_{ba}}{N} + O(\frac{\rho_{ij}}{(N |\eta_i|)^{3/2}}) \right)
\]
\[
\lesssim N^{-5/2} \frac{1}{N |\eta_j|} + N^{-5/2} N^2 \sqrt{\frac{\rho_1}{N |\eta_j|}} \left( \frac{\rho_{ij}}{N |\eta_i|^{3/2}} \right) \Psi_{(1, i, j)}
\]
and similarly
\[
|\Xi_2(\{i\}, \{j\})| = O^2 \left( \frac{\Psi_{(1, i, j)}}{L^{1/2}} \right).
\]
This concludes the proof of (107).

Proof of (112). The proof of (112) is very similar to that of (107) and we again consider the cases (i) \(k_1\) odd, \(k \geq 4\), (ii) \(k_1\) even, \(k \geq 3\), (iii) \(k_3 = 3, k_1 = 1\), (iv) \(k = 3, k_1 = 3\) separately.

Case (i), \(k_1\) odd, \(k \geq 4\). In this case we estimate \(|\partial_{a_1} (G_1)_{ba}| \prec 1\) and obtain from (116) that
\[
|\Phi_k(P_Y)| \prec \rho_1^{-1} N^{-(k+3)/2} N^2 \Psi_{S(P_Y)} \lesssim N^{-(k-3)/2} \Psi_{(1) \cup S(P_Y)},
\]
from \(N^{-1} \lesssim \rho_1 N^{-1} |\eta_i|^{-1}\), confirming (112).

Case (ii), \(k_1\) even, \(k \geq 3\). Since \(k_1\) is odd it follows by parity and the local law that \(|\partial_{a_1} (G_1)_{ba}| \prec |(G_1)_{ba}|\) and therefore, by (116) and a Ward-estimate it follows that
\[
|\Phi_k(P_Y)| \prec \rho_1^{-1} N^{-(k+3)/2} N^2 \sqrt{\frac{\rho_1}{N |\eta_i|}} \Psi_{S(P_Y)} \lesssim N^{-(k-2)/2} \Psi_{(1) \cup S(P_Y)},
\]
confirming (112).

Case \(k = 3, k_1 = 3\). The derivatives acting on \((G_1)_{ab}\) results in four \(G_1\) factors with a total of four \(a\) and four \(b\) indices and we obtain
\[
\Phi_3(\emptyset) = - \kappa_4 m_1^3 N^{-3} \sum_{ab} (G_1)_{ba}^3 (G_1)_{ba} \Phi_3(\emptyset) + O \left( \rho_1^{-1} N^{-3} \sum_{ab} |(G_1)_{ab}| \right)
\]
\[
= - \kappa_4 m_1^3 m_3 + O \left( \frac{1}{N^{3/2} \sqrt{\eta_i}} \right) = - \kappa_4 m_1^3 m_3 + O \left( \frac{\Psi_{(1)}}{L} \right),
\]
which gives one of the leading terms in (112).

Case (iii), \(k = 3, k_1 = 1\). For \(k_1 = 1\) the derivative of \((G_1)_{ba}\) is given by
\[
\sum_{a_1} \partial_{a_1} (G_1)_{ba} = -(G_1)_{ba} - (G_1)_{ba} (G_1)_{ba} = -m_1^2 + \Omega \left( \frac{\rho_1}{N |\eta_i|} \right)
\]
If \(P_Y = \{i, i\}\) for some \(i \in \{1, r\}\), then we obtain from (117) that
\[
\partial_{ab, ba} Y_i = \partial_{ba, ab} Y_i = m_1 m_3 \frac{2}{N} + O \left( \frac{\rho_1^{1/2}}{(N |\eta_i|)^{3/2}} \right),
\]
while both the \(\partial_{ab,ab}\) and \(\partial_{ba,ba}\) derivatives lead to delta functions \(\delta_{ab}\) and are therefore lower order, and thus

\[
\Phi_3\{i, i, \emptyset\} = -\frac{2}{N^2} m_1 m_i m_i' m_i'' \kappa(ab, ba, ab, ba) \frac{2}{N^2} + O_\prec\left(\frac{\Psi_{(1, 1)}}{L^{1/2}}\right)
\]

\[
= -\frac{2\kappa_3}{N^2} m_1 m_i m_i' + O_\prec\left(\frac{\Psi_{(1, 1)}}{L^{1/2}}\right) = -\frac{\kappa_3}{2N^2} (m_i') (m_i'') + O_\prec\left(\frac{\Psi_{(1, 1)}}{L^{1/2}}\right),
\]

giving the other leading term in (112). On the other hand, if \(P_{Y} = \{i, j\}\) for some \(1 < i < j\), then we obtain a \(\delta_{ab}\) from (117) and therefore the leading term is at most of size \(\Psi_{(1, 1, i)} L^{-1/2}\) and we obtain

\[
|\Phi_3\{i, j, \emptyset\}| = O_\prec\left(\frac{\Psi_{(1, 1, j)}}{L^{1/2}}\right).
\]

This concludes the proof of (112) for the case \(k_1 = 1, k = 3\).

Case (iv), \(k = 2\). In case \(k = 2\) there are four sub-cases to consider; \(k_1 = 2, P_{Y} = \{i, i\}, k_1 = 1\) or \(P_{Y} = \{i, j\}\) for \(i \neq j\).

If \(k_1 = 2\), then the derivative is given by \(\partial^2 (G_1)_{ab} = (G_1)^3\) with three \(a\) and three \(b\) indices, so that by parity either all three matrices have indices \(ab, ba\), or only one with the remaining two having \(aa, bb\). For all three matrices having \(ab, ba\) indices, we can gain two factors of \(\sqrt{\rho_1 / N \eta_1}\) via Ward-estimates over the naive size \(N^{-1/2}\) in order to obtain \(N^{-3/2} \eta_1^{-1} \leq \Psi_{(1, 1)} L^{-1/2}\). If two matrices have indices \(aa, bb\), then we estimate

\[
\rho_1^{-1} N^{-5/2} \sum_{ab} (G_1)_{ba} (G_1)_{aa} (G_1)_{ba} = m_1 N^{-5/2} \sum_a (G_1)_{aa} (G_1)_{1a} + O_\prec\left(\frac{\rho_1}{N^{3/2} \eta_1}\right)
\]

\[
= O_\prec\left(\frac{\rho_1^{1/2}}{N^{3/2} \eta_1^{1/2}}\right) = O_\prec\left(\frac{\Psi_{(1)}}{L^{1/2}}\right),
\]

in order to conclude

\[
|\Phi_2(\emptyset)| \prec \frac{\Psi_{(1)}}{L^{1/2}}.
\]

Next, if \(P_{Y} = \{i, i\}\), then we obtain from (117) that

\[
\Phi_2(\{i, i\}) = \frac{m_1'}{m_1} \sum_{ab} \kappa(ab, ba, ab, ba) \frac{2}{N} + O_\prec\left(\frac{\Psi_{(1, 1)}}{L^{1/2}}\right)
\]

\[
= \frac{m_1'}{m_1} \frac{2\kappa_3}{N^{1/2}} (1, G_1 1) + O_\prec\left(\frac{\Psi_{(1, 1)}}{L^{1/2}}\right) = O_\prec\left(\frac{\Psi_{(1, 1)}}{L^{1/2}}\right).
\]

The case \(k_1 = 1\) can be estimated by

\[
|\Phi_2(\{i\})| \lesssim \frac{N^{-7/2}}{\rho_1 \rho_i} \sum_{ab} \left| (G_1)_{ba} (G_1)_{aa} (G_1)_{ba}^2 (G_1)_{ba} \right|
\]

\[
\lesssim \frac{N^{-7/2}}{\rho_1 \rho_i} \sum_a \left| (G_1^2)_{1a} \right| + \frac{N^{-7/2}}{\rho_1 \rho_i} \sum_{ab} \sqrt{\frac{\rho_1}{N \eta_1}} \left| (G_1^2)_{ba} \right| \lesssim \frac{\rho_1^{1/2}}{N^{3/2} \eta_1^{1/2}} \frac{\rho_i}{N^{3/2} \eta_1^{1/2}} \lesssim \frac{\Psi_{(1, 1)}}{L^{1/2}}.
\]

For the final case \(P_{Y} = \{i, j\}\) with \(i \neq j\) we obtain from (117) that

\[
|\Phi_2(\{i, j\})| \lesssim \frac{N^{-7/2}}{\rho_1 \rho_i \rho_j} \sum_{ab} \left| (G_1)_{ba} \right| \left| \delta_{ba} \right| + O_\prec\left(\frac{\rho_1^{1/2}}{N \eta_1^{3/2}}\right) \left| \delta_{ba} \right| + O_\prec\left(\frac{\rho_j^{1/2}}{(N \eta_1)^{3/2}}\right)
\]

\[
\lesssim \frac{\Psi_{(1, 1, j)}}{L^{1/2}}.
\]

This concludes the proof of (112).  \(\square\)
5. Functional CLT: Proof of Theorems 2.3–2.4.

In this section we prove our main results, the functional central limit theorems using the resolvent CLT, Theorem 4.1. Via standard representation formulas this involves fairly standard but tedious calculations. We first give a detailed calculation for the case sharp cut-off case in Section 5.1 using the less-known Pleijel’s formula which proves Theorems 2.3. The proof of Theorem 2.4 in Section 5.2 relies on similar calculations using the more conventional Helffer-Sjöstrand formula; the details are deferred to Appendix E.

5.1. Proof of the functional CLT for the sharp cut-off.

Proof of Theorem 2.3. Let $\hat{A} := A - \langle A \rangle$, and define

$$1_{K,i_0}(i) := 1(|i - i_0| \leq K).$$

We recall the rigidity bound (see e.g. [21, Lemma 7.1, Theorem 7.6] or [25, Section 5]):

$$|\lambda_i - \gamma_i| \lesssim \frac{1}{N^{2/3}K^{1/3}},$$

where $i := i \wedge (N + 1 - i)$. Here $\gamma_i$ are the classical eigenvalue locations (quantiles) defined by

$$\int_{\gamma_i}^{\gamma_i + 1} \rho(x) \, dx = \frac{i}{N}, \quad i \in [N],$$

where we recall $\rho(x) = \rho_m(x) = (2\pi)^{-1/2} \sqrt{(4 - x^2)^+}$. We now present the proof in the bulk regime, the edge is completely analogous and so omitted. Define $\eta_{K}(\gamma_{i_0})$ implicitly by

$$\eta_{i_0} = \eta_{K}(\gamma_{i_0}) := \frac{K}{N \rho(\gamma_{i_0} + \eta_{K}(\gamma_{i_0}))}.$$

Then, by (122) and (123), we readily conclude

$$\sqrt{\frac{N}{2K}} \sum_{|i - i_0| \leq K} \langle u_i, \hat{A} u_i \rangle = \sqrt{\frac{N}{2K}} \sum_{i = 1}^{N} 1_{K,i_0}(i) \langle u_i, \hat{A} u_i \rangle = \frac{N^{3/2}}{\sqrt{2K}} \langle P(W) \hat{A} \rangle + O_{\prec} \left( \frac{1}{\sqrt{K}} \right),$$

where we defined the spectral projection

$$P(W) = 1(\gamma_{i_0} - \eta_{i_0} \leq W \leq \gamma_{i_0} + \eta_{i_0}),$$

and used that $|\langle u_i, \hat{A} u_i \rangle| \lesssim N^{-1/2}$ by [25, Theorem 1].

Using Pleijel’s representation formula of the spectral projection of a Hermitian matrix in terms of contour integral of its resolvent in [23, Eq. (13)] (see also [47, Eq. (5)]), we find that (see Appendix F for more details)

$$\frac{N^{3/2}}{\sqrt{2K}} \langle P(W) \hat{A} \rangle = \frac{N^{3/2}}{2\pi i \sqrt{2K}} \int_{\Gamma_{K,i_0}} \langle G(z) \hat{A} \rangle \, dz + O_{\prec} \left( \frac{N \eta_\delta}{\sqrt{K}} \right),$$

with $\Gamma_{K,i_0}$ the contour oriented counter-clockwise and defined by

$$\Gamma_{K,i_0} := \{ z \in \mathbb{C} \mid \Re z \in [\gamma_{i_0} - \eta_{i_0}, \gamma_{i_0} + \eta_{i_0}] \text{ and } |\Im z| = M \}$$

$$\cup \{ z \in \mathbb{C} \mid \Re z \in [\gamma_{i_0} - \eta_{i_0}, \gamma_{i_0} + \eta_{i_0}] \text{ and } |\Im z| \in [\eta_\delta, M] \},$$

for any $M > 0$, and $\eta_\delta$ such that $N^{-1} \ll \eta_\delta \ll K^{1/2}/N$.

By Young’s inequality, for any $p \geq 2$ and $\delta > 0$, we get from (128) that

$$E \left| \frac{N^{3/2}}{\sqrt{2K}} \langle P(W) \hat{A} \rangle \right|^p = \left( (1 + O(N^{-\delta})) E \left| \frac{N^{3/2}}{\sqrt{2K}} \int_{\Gamma_{K,i_0}} \langle G(z) \hat{A} \rangle \, dz \right|^p \right)^{\frac{1}{p}} + O \left( \left( \frac{N^{1+\delta} \eta_\delta}{\sqrt{K}} \right)^{\frac{1}{p}} \right) \cdot \left( (1 + O(N^{-\delta})) \right)^{\frac{1}{p}}.$$
By (91), it follows that for even \( p \) (for odd \( p \) the leading term is zero) we have
\[
E \left[ \frac{N^{3/2}}{\sqrt{2K}} \langle P(W) A \rangle^p \right] = \left( (1 + O(N^{-\delta})) \right) \left( \frac{N}{2K} \right)^{p/2} \sum_{P \in \text{Pair}([p])} \prod_{i,j} \frac{1}{4\pi^2} \int_{\Gamma_{K,i_0}} dz_i \int_{\Gamma_{K,i_0}} dz_j \\
\times \left( \frac{m_i^2 m_j^2 (A^2)}{1 - m_i m_j} + \sigma m_i^2 m_j^2 (A A^1) \right) \\
+ \kappa_3 m_i^2 m_j^2 (a_i a_j) \\
+ O \left( \frac{N^4}{\sqrt{NNM}} \left[ (\frac{MN}{K})^{p/2} + (\frac{MN}{K})^{-p/2} \right] \right).
\]  
(131)
for any \( \xi, \delta > 0 \). In estimating the error term coming from (91) we used that
\[
\left( \prod_{i \in [p]} \int_{\Gamma_{K,i_0}} dz_i \right) \frac{1}{\sqrt{NN\eta_0}} \prod_{i \in [p]} \frac{N^{\xi}}{N^{\xi}} \leq \frac{N^{\xi}}{\sqrt{NNM}} \left[ (\frac{MN}{K})^{p/2} + (\frac{MN}{K})^{-p/2} \right].
\]
where \((MN/K)^{p/2}\) comes from the vertical lines of the contour \( \Gamma_{K,i_0} \) and \((MN/K)^{-p/2}\) from the horizontal ones. Note that in order to apply (91) we had to choose \( \eta_0 \gg N^{-1} \), which ensures \( L = N \min_i \langle \{ \eta_i \} | \rho_i \rangle \sim N\eta_0 \gg 1 \) (since we are in the bulk regime), with \( \eta_i := \min_i |\eta_i| \). It is clear that we can choose \( \xi, \delta \), and \( \eta_0 \ll K^{1/2}/N, M \ll K/N \) so that the error term in (131) is bounded by \( N^{-c(p,\epsilon)} \) for some constant \( c(p,\epsilon) > 0 \), where \( N^{1/4} \leq K \leq N^{1/4} \). In the following \( \eta_0 \ll M \ll K/N \) ensures that only the horizontal lines of \( \Gamma_{K,i_0} \) contribute to the integral, the vertical lines are negligible giving a contribution \( MN/K \).

We start computing
\[
\frac{2N}{K} \int_{\gamma_0 - \eta_0}^{\gamma_0 + \eta_0} dx dy \left\{ \frac{\sigma m^2 m^2}{1 - \sigma m^2 m^2} - \frac{\sigma m^2 m^2}{1 - \sigma m^2 m^2} \right\} \\
= -\frac{N}{K} \int_{\gamma_0 - \eta_0}^{\gamma_0 + \eta_0} \sqrt{(4-x^2)_+ (4-y^2)_+} \, dx \, dy \\
+ 1(\sigma = \pm 1) \frac{2N\pi}{K} \int_{\gamma_0 - \eta_0}^{\gamma_0 + \eta_0} \sqrt{(4-x^2)_+ (4-y^2)_+} \, dx \, dy \\
- \frac{N}{K} \int_{\gamma_0 - \eta_0}^{\gamma_0 + \eta_0} \frac{(1 - \sigma^2) \sqrt{(4-x^2)_+ (4-y^2)_+}}{\sqrt{x^2 + y^2} + (1 - \sigma^2) - xy \sigma (1 + \sigma^2)} + O(\sqrt{M}).
\]  
(132)
In particular, for \( K \ll N \) we get that the rhs. of (132) is equal to
\[
8\pi^2 1(\sigma = 1) + 1(\sigma = -1) \frac{2N\pi}{K} \int_{I_{\eta_0}} \sqrt{4-x^2} \, dx + O\left( \frac{\sqrt{M} + K}{N} \right),
\]
with \( I_{\eta_0} := [\gamma_0 - \eta_0, \gamma_0 + \eta_0] \cap [-\gamma_0 - \eta_0, -\gamma_0 + \eta_0] \). Note that for \( \eta_0 = [cN] \) we have
\[
\frac{2N\pi}{K} \int_{I_{\eta_0}} \sqrt{4-x^2} \, dx = \frac{4N\pi}{K} |I_{\eta_0}| + O\left( \frac{K}{N} \right).
\]
We now distinguish two cases: (i) \( c \neq 1/2 \), (ii) \( c = 1/2 \). If \( c \neq 1/2 \) then \( |\gamma_0| \geq |c - 1/2| \) and so \( I_{\eta_0} \) is empty, i.e.
\[
\frac{2N\pi}{K} \int_{I_{\eta_0}} \sqrt{4-x^2} \, dx = 0.
\]
On the other hand, for \( c = 1/2 \) we have
\[
\frac{2N\pi}{K} \int_{I_{\eta_0}} \sqrt{4-x^2} \, dx = \frac{4N\pi}{K} |I_{\eta_0}| + O\left( \frac{K}{N} \right) = 8\pi^2 + O\left( \frac{K}{N} \right),
\]
where we used that for \( c = 1/2 \) we have \( |I_{0_0}| = 2\eta_0 + \mathcal{O}(N^{-1}) \), as a consequence of \( \gamma_{0_0} = \mathcal{O}(N^{-1}) \), and that \( \rho(i\eta_0) = \pi^{-1} + \mathcal{O}(\eta_0) \), with \( \eta_0 \lesssim Kn^{-1} \) in the bulk. By analogous computations we conclude that in the edge regime the r.h.s. of (132) is equal to \( 1(\sigma = 1)8\sqrt{2\pi^2}/3 \).

For mesoscopic scales the third and of the fourth term are negligible. This concludes the proof of Theorem 2.3.

\[ \square \]

5.2. **Proof of the functional CLT for smooth cut-off.**

**Proof of Theorem 2.4.** For simplicity we present the proof in the macroscopic scale and in the mesoscopic scale in the bulk. The computation of the leading term and the estimate of the error terms at the edge are completely analogous and so omitted.

For any \( z = x + i\eta \in \mathbb{C} \) we define the almost analytic extension of \( f \in H^2 \) by

\[
f_c(z) = f_c(x + i\eta) := [f(x) + i\eta \partial_x f(x)] \chi(N^\alpha \eta),
\]

where \( \chi \) is a smooth cut-off equal to 1 for \( \eta \in [-5, 5] \) and equal to 0 for \( \eta \in [-10, 10]^\circ \). Note that

\[
|\partial_x f_c| \lesssim N^{2\alpha} |g''| |\eta| + N^\alpha (|g| + N^\alpha |g'|) |\chi'|,
\]

where \( 2\partial f_c := \partial x + i\partial \eta \).

By the Helffer-Sjöstrand formula we have that

\[
f_c(z) = \frac{1}{\pi} \int_{\mathbb{C}} \frac{\partial_x f_c(z)}{z - \lambda} \, d^2z = \frac{2}{\pi} \Re \int_{\mathbb{R}^+} \int_{\mathbb{R}_+} \frac{\partial_x f_c(z)}{z - \lambda} \, d\eta \, dx,
\]

where \( d^2z = dx \, d\eta \) is the Lebesgue measure on \( \mathbb{R}^2 \). We recall the following notation

\[
L_N(f, I) = \sum_{i=1}^{N} f(\lambda_i) - E \sum_{i=1}^{N} f(\lambda_i)
\]

(136)

\[
L_N(f, \hat{A}) := L_N(f, \hat{A}_d) + L_N(f, A_{od}) = \sum_{i=1}^{N} f(\lambda_i) (u_i, \hat{A} u_i).
\]

(137)

Using (135), we write

\[
L_N(f, I) = \frac{2N}{\pi} \Re \int_{\mathbb{R}} \int_{\mathbb{R}_+} \partial_x f_c(z) \langle G(x + i\eta) - E G(x + i\eta) \rangle \, d\eta \, dx
\]

(138)

\[
L_N(f, \hat{A}) = \frac{2N^{1+\alpha/2}}{\pi} \Re \int_{\mathbb{R}} \int_{\mathbb{R}_+} \partial_x f_c(z) \langle G(x + i\eta) \hat{A} \rangle \, d\eta \, dx.
\]

Using that \( |(u_i, \hat{A} u_i)| \lesssim N^{-1/2} \), writing \( G \) in spectral decomposition, we conclude

\[
|\langle G(x + i\eta) \hat{A} \rangle| \leq \frac{1}{N} \sum_{i=1}^{N} \frac{|(u_i, \hat{A} u_i)|}{|\lambda_i - z|} \ll \frac{1}{\sqrt{N}} \left( 1 + \frac{1}{N\eta} \right)
\]

(139)

for any \( \eta \geq N^{-100} \), where we used that the local law for \( |(G - E G)| \ll (N\eta)^{-1} \) holds for any \( \eta \geq N^{-100} \) (e.g. see [14, Appendix A]).

Then, using (139), (134), and the local law \( |(G - E G)| \ll (N\eta)^{-1} \), we readily conclude that

\[
L_N(f, I) = \frac{2N}{\pi} \Re \int_{\mathbb{R}} \int_{\eta_0}^{\eta_0} \partial_x f_c(z) \langle G(x + i\eta) - E G(x + i\eta) \rangle \, d\eta \, dx + O_c(\eta_0 N^\alpha)
\]

(140)

\[
N^{\alpha/2} L_N(f, \hat{A}) = \frac{2N^{1+\alpha/2}}{\pi} \Re \int_{\mathbb{R}} \int_{\eta_0}^{\eta_0} \partial_x f_c(z) \langle G(x + i\eta) \hat{A} \rangle \, d\eta \, dx + O_c(\eta_0^2 N^{(1+3\alpha)/2}),
\]

where we defined

\[
\eta_0 := N^{-1+\epsilon}, \quad \eta_a := 10N^{-a},
\]
for some small $\epsilon > 0$ such that $\eta_0 \ll \eta_a$. Note that in (40) we used that $\chi'((N^3/\eta) = 0$ on $\eta \in [0, \eta_0]$ since $\eta_0 \ll N^{-6}$, and so that $|\partial_x f_{\epsilon}c| \leq N^{2a}\eta$ by (34). We remark that the regime $\eta \leq N^{-100}$ in (40) is bounded trivially by $N^{-100+2a}$ using that $|f(GA)| \leq \eta^{-1}$ and that $|\partial_x f_{\epsilon}c| \leq N^{2a}\eta$.

With the formulas (40) we thus reduced the proof of the functional CLT for general test function $f$ to the CLT for resolvents as given in Theorem 4.1, modulo detailed calculations of the leading terms. These calculations are deferred to Appendix E and with their help we conclude the proof of Theorem 2.4. $\square$

Appendix A. Case of vanishing variances in Theorem 2.4

In this section we give a short explanation for the cases of vanishing variances in Theorem 2.4 as listed in Remark 2.8.

The fact that for constant $f$ the limiting processes vanish is obvious since in this case $\text{Tr } f(W)A$ is deterministic. Similarly, for linear $f(x) = bx$ and $w_2 = 0$ the diagonal processes $\xi_{\alpha}$ vanish since $\text{Tr } f(W)A_{\alpha} = b \text{Tr } W A_{\alpha} = 0$ almost surely if $w_2 = 0$. For the case of quadratic $f(x) = cx^2$ and $\kappa_4 = -1 - \sigma^2$, i.e. $|w_{12}| = 1/\sqrt{N}$ almost surely, we have

$$\text{Tr } W^2 A_{\alpha} = \sum_a (A_{\alpha})_{aa} \left( \sum_b w_{ab} w_{ba} \right) = \sum_a (A_{\alpha})_{aa} \left( \frac{N - 1}{N} + w_{aa}^2 \right)$$

almost surely, so that $\text{Var} (\text{Tr } W^2 A_{\alpha}) \lesssim (|A_{\alpha}|^2)/N$. For $\sigma = 1$ and skew-symmetric $A_{\alpha\beta} = -A_{\beta\alpha}$ is clear that $2 \text{Tr } f(W)A_{\alpha\beta} = \text{Tr } f(W)A_{\alpha\beta} + \text{Tr } (f(W)A_{\alpha\beta})^T = \text{Tr } f(W)(A_{\alpha\beta} + A_{\beta\alpha}^T) = 0$ due to $W$. $f(W)$ being almost surely real-symmetric.

It remains to consider the cases of vanishing variances for $\sigma = -1$, i.e. when $W = D + iR$ for some real diagonal $D$ and some real skew-symmetric $R = -R^T$. If $D = 0$, then by the exact symmetry of the spectrum, $\lambda_i = -\lambda_{N+1-i}$ and $u_i = \overline{u}_{N+1-i}$ (up to phase), we immediately see that all three linear statistics are constant for odd functions $f$. In case $D \neq 0$ the variance is not algebraically zero but it is vanishing for large $N$. To illustrate this mechanism, we consider the odd function $\phi(x) = x^3$. Then (a)–(c) in Remark 2.8 are saying that $\text{Tr } (W^3 - 3W), \text{Tr } (W^3 - 2W)A_{\alpha\beta}$ and $\text{Tr } W^3 A_{\alpha\beta}$ fluctuate on a scale $\ll 1$. Indeed,

$$\text{Tr } W^3 A = \text{Tr } D^3 A + i \text{Tr } (D^2 R + D R D + R D^2) A - \text{Tr } (D R^2 + R D R + R^2 D) A - i \text{Tr } R^3 A$$

so that

$$\text{Tr } (W^3 - 3W) = \text{Tr } D^3 - 3 \text{Tr } D (1 + R^2)$$

is $\ll 1$ since $(R^2)_{aa} = -\sum_b R_{ab}^2 \approx 1$. Similarly,

$$\text{Tr } (W^3 - 2W)A_{\alpha\beta} = \text{Tr } D^3 A_{\alpha\beta} - \text{Tr } (2D (1 + R^2) + R D R) A_{\alpha\beta}$$

since $\text{Tr } R^3 A_{\alpha\beta} = 0$ due to $R = -R^T$ and $A_{\alpha\beta} = A_{\beta\alpha}$ and the rhs. is $\ll 1$ since

$$|\text{Tr } R D R A_{\alpha\beta}| = \left| \sum_{ab} R_{ab}^2 D_{ab} (A_{\alpha\beta})_{aa} \right| \leq N^{-1/2} (|A_{\alpha\beta}|^2).$$

Finally,

$$\text{Tr } W^3 A_{\alpha\beta} = i \text{Tr } (D^2 R + D R D + R D^2) A_{\alpha\beta} - \text{Tr } (D R^2 + R D R + R^2 D) A_{\alpha\beta} - i \text{Tr } R^3 A_{\alpha\beta}$$

is $\ll N^{-1/2} \sqrt{A_{\alpha\beta} A_{\alpha\beta}^T}$ due to $A_{\alpha\beta}$ being off-diagonal. A similar argument works for any odd polynomial.

Appendix B. Proofs of Lemmata 3.7–3.8

Proof of Lemma 3.7. We only prove (75), the proof of (76) is analogous and so omitted. In order to prove (75) we will often use the resolvent identity

$$G(z_i) = R(z_i) - R(z_i)DG(z_i).$$

(141)

In particular, using (144) repeatedly, we find that

$$\langle G_1 G_2 \rangle = \langle R_1 R_2^\dagger \rangle - \langle R_1 D R_1 R_2^\dagger \rangle - \langle R_1 R_2^\dagger D R_2^\dagger \rangle + \langle R_1 D R_1 D R_1 R_2^\dagger \rangle + \ldots$$

(142)
where we used the short-hand notation $G_i = G(z_i), R_i = R(z_i)$. First we prove that we can stop the expansion \((42)\) after $q = 10^{-4}$ resolvent identities for $G_1$ and $G_2$, keeping the last resolvent as $G_i$, at the price of negligible error smaller than $N^{-5}$ with very high probability. We now bound a representative term, with $l$ factors $R_1$ and $m$ factors $R_2$ and $l + m = q$, to explain how the expansion in \((42)\) is truncated. By a Schwarz inequality we have

$$\begin{align*}
|\langle R_1 DR_1 \ldots R_1 D G_1 R_2 DR_2 \ldots R_2 D G_2 \rangle| &\leq \langle R_1 DR_1 \ldots R_1 D G_1 \rangle^{1/2} \langle R_2 DR_2 \ldots R_2 D G_2 \rangle^{1/2} \\
&\leq \frac{1}{\eta_1^2 \eta_2^2} \langle 3 R_1 DR_1 \ldots R_1 D 3 R_1 D R_2^* \ldots R_2^* D \rangle^{1/2} \langle 3 R_2 DR_2 \ldots R_2 D 3 R_2 D R_2^* \ldots R_2^* D \rangle^{1/2}.
\end{align*}$$

Here we estimated $\|G_i^*\| \leq \eta_i^{-2}$ and used the Ward identity $R_i R_i^* = \eta_i^{-1} \Im R_i$. Note that to go from the first to the second line we used that $\langle R_1^2 D R_2^* \ldots (R_2^*)^d D (R_2^*)^e \rangle = \langle R_2 D R_2 \ldots R_2^* D R_2^* \rangle$ by cyclicity of the trace. In the following, to bound the terms in the rhs. of \((43)\) we write $2i \Im R_i = R_i - R_i^*$, since we do not need to exploit any additional gain from $3 R_i$. For simplicity we assume that $l$ is even and denote $n = 4(l - 1)p$, then, denoting by $E_D$ the expectation with respect to the diagonal randomness $D$ and using that the $p$-th moment of the entries of $D$ are bounded by $C_p N^{-p/2}$ by \((5)\), we estimate

$$\begin{align*}
E_D \langle |R_1 DR_1 \ldots R_1 D R_1^* \ldots R_2^* D |^{2p} &\leq E_D \frac{1}{N^{2p}} \sum_{i_1, \ldots, i_n \in [N]} D_{i_1 i_1} \ldots D_{i_n i_n} (R_1)_{ab} (R_1)_{cd} \ldots (R_1)_{vw} \\
&\leq \frac{1}{N^{2p}} \frac{1}{N^{2(l-1)p}} N^{2(l-1)p} \left( \frac{1}{N \eta_1} \right)^{p(l-2)} \frac{1}{N \eta_1} \left( \frac{1}{N \eta_1} \right)^{p(l-2)} = \frac{1}{N^{2p}} \frac{1}{N \eta_1} \left( \frac{1}{N \eta_1} \right)^{p(l-2)},
\end{align*}$$

(44)

where the indices $a, b, c, \ldots, v, w$ are from the set \{$i_1, i_2, \ldots, i_n$\}, their precise allocation is irrelevant. In particular, in the second line of \((44)\) we neglected the fact the the summations involve $2p$ traces and wrote directly the summation from $i_1$ to $i_n$, since the cyclic structure does not play any role in the estimate. The factors $N^{-2(l-1)p}$ and $N^{2(l-1)p}$ in \((44)\) come from the pairings in the expectations $E_D$ and from the number of effective summations, respectively. More precisely, in \((44)\) we used that the main contribution comes from the case when a $D_{i_1 i_1}$ is paired with some other $D_{i_n i_n}$, since in this way only half of the summations collapse. For higher moments the effective summation contains even less indices. The factor $(N \eta_1)^{-p(l-2)}$ comes from the fact that after the pairings at least $p(l-2)$ factors $R_1$ are off-diagonal gaining a factor $(N \eta_1)^{-1/2}$ for each one of them by the local law $|\langle R_1 \rangle_{ab} | \leq (N \eta_1)^{-1/2}$. Combining \((43)\) and \((44)\) for any $p \in \mathbb{N}$, we conclude that

$$|\langle R_1 DR_1 \ldots R_1 D G_1 R_2^* DR_2^* \ldots R_2^* D G_2^* \rangle| \prec \frac{1}{N \eta_1 \eta_2} \frac{1}{(N \eta_1)^{q-4}/2} \leq N^{-5},$$

since $\eta_1, \eta_2 \geq N^{-1+\epsilon}$ and $q = 10^{\epsilon^{-1}}$. This implies that the expansion in \((42)\) can be stopped after a finite number of terms.

In order to conclude the proof we have to estimate the fully expanded terms in \((42)\), i.e. the terms which involve only $R_1, R_2$ and no $G_1, G_2$ appear. We now give two different bounds for the terms in \((42)\): the first one in \((48)\) is better in the regime $|z_1 + z_2| \geq \eta^*$, the second one in \((49)\) is better in the opposite regime $|z_1 + z_2| < \eta^*$. To bound these terms we use the following strategy: Step (i) for each factor $R_1 R_2^*$ first perform a resolvent identity:

$$R_1 R_2^* = R(z_1) R(-z_2) = -R(z_1) R(-z_2) \leq \frac{R(z_1) - R(-z_2)}{z_1 + z_2};$$

(45)

Step (ii) compute high moments with respect to $E_D$. In particular, in each term in the expansion \((42)\) we can perform one resolvent identity if in the trace only one $R_1$ or one $R_2$ appears, and two resolvent identities otherwise. To make this argument clearer we bound explicitly a representative term, all the
other terms are bounded exactly in the same way and so omitted. Using (143), we start with the bound
\[ E_D \left| \langle R_1 DR_1 DR_1 R_2^\dagger \rangle \right|^{2p} \lesssim \frac{1}{|z_1 + z_2|^{2p}} E_D \left| \langle (R_1) DR_1 DR_1 (z_1) \rangle \right|^{2p} + \left| \langle (R_1) DR_1 DR_1 (z_2) \rangle \right|^{2p}. \] (146)

Note that transposes disappeared in the rhs. of (146), and in the following we do not make distinction between \( R_1 \) and \( R_2 \), or their adjoints, every term is bounded in terms of \( \eta_* = \eta_1 \land \eta_2 \). The bound of the two terms in the rhs. of (146) is analogous and so we only consider the second one
\[ E_D \left| \langle R_1 DR_1 DR_1 DR_1 R_2^\dagger \rangle \right|^{2p} \]
\[ = \frac{1}{N^{2p}} E_D \sum_{i_1, \ldots, i_{4p} \in [N]} D_{i_1 i_1} \ldots D_{i_{4p} i_{4p}} R_{ab} (RR)_{cd} \ldots R_{tu} (RR)_{vw} \]
\[ \lesssim \frac{1}{N^{2p}} \frac{1}{N^{2p}} \sum_{a_1, \ldots, a_{4p} \in [1, \ldots, 4p], a_1, \ldots, a_{4p} \in [N]^{2p}} R_{a_1 a_2} (RR)_{a_3 a_4} \ldots R_{a_2p - 3a_2p - 2} (RR)_{a_2p - 1 a_2p} \]
\[ \lesssim \frac{1}{(N\eta_*)^{2p}}, \] (147)
where the indices \( a, b, c, d, \ldots, t, u, v, w \) are from the set \( \{ i_1, i_2, \ldots, i_{4p} \} \), their precise allocation is irrelevant, and we used the notation \( R \in \{ R(z_1), R(-z_2) \} \). Additionally, in (147) we used that, as in (144), the leading order contribution to \( E_D \) comes from the pairings and so that from the second to the third line of (147) 2p indices collapse. To go from the third to the fourth line we used that there are exactly 2p factors \( (RR)_{ab} \), and that \( |(RR)_{ab}| \approx \eta_*^{-1} \) and \( |R_{ab}| \approx 1 \). Finally, combining (146) and (147), we conclude that
\[ |\langle R_1 DR_1 DR_1 R_2^\dagger \rangle| \lesssim \frac{1}{N\eta_* |z_1 + z_2|}. \] (148)

Using that an analogous bound holds for all the other terms in the expansion (142), this concludes the proof of (73).

We now prove another bound for the terms (142) which improves (148) in the regime \(|z_1 + z_2| < \eta^*\). Similarly to (146), also in this case we bound a representative term of the expansion (142):
\[ E_D \left| \langle R_1 DR_1 DR_1 R_2^\dagger \rangle \right|^{2p} \]
\[ = \frac{1}{N^{2p}} E_D \sum_{i_1, \ldots, i_{4p} \in [N]} D_{i_1 i_1} \ldots D_{i_{4p} i_{4p}} (R_1)_{ab} (R_1 R_1 R_1)_{cd} \ldots (R_1)_{tu} (R_1 R_1 R_1)_{vw} \]
\[ \lesssim \frac{1}{N^{2p}} \sum_{a_1, \ldots, a_{4p} \in \{ 1, \ldots, 4p \}, \{ i_1, \ldots, i_{2p} \} \subset [N]^{2p}} (R_1)_{a_1 a_2} (R_1 R_1 R_1)_{a_3 a_4} \ldots (R_1)_{a_2p - 3a_2p - 2} (R_1 R_1 R_1)_{a_2p - 1 a_2p} \]
\[ \lesssim \frac{1}{(N\eta_1 \eta_2)^{2p}}, \] (149)
where the indices \( a, b, c, d, \ldots, t, u, v, w \) are as in (147), their precise allocation is irrelevant. Additionally, as in (147) we used that the leading order contribution to \( E_D \) comes from the pairings so that from the second to the third line of (149) 2p indices collapse. To go from the third to the fourth line we used that there are exactly 2p factors \( (R_1 R_1 R_1)_{ab} \), and that \( |(R_1 R_1 R_1)_{ab}| \approx (\eta_1 \eta_2)^{-1} \) and \( |(R_1)_{ab}| \approx 1 \).

**Proof of Lemma 5.8.** Since both the l.h.s. and r.h.s. of (80) are symmetric in \( \eta_1 \) and \( \eta_2 \), without loss of generality we assume that \( \eta_* = \eta_1 \). Recalling that \( L = N (pn)_* \), with \( (pn)_* := (p_1 \eta_1) \land (p_2 \eta_2) \), it readily follows that
\[ \frac{1}{N \eta^2 |z_1 - z_2|} \lesssim \frac{\rho_1}{L \eta_1 \eta_2}, \]
where we used that \(|z_1 - z_2| \gtrsim \eta_1 + \eta_2 \). Hence if \( \rho_* = \rho_1 \) there is nothing else to prove. In the remainder of the proof we assume that \( \rho_* = \rho_2 \). Using the definition of \( L \), we readily see that (80) is
The proof of \((\eta \rho) \cdot \eta_2 \lesssim \rho_2 |z_1 - z_2| \eta_1\). \hfill (150)

We now distinguish two cases: (i) \(\rho_2 \leq |z_1 - z_2|^{1/2}\), (ii) \(\rho_2 > |z_1 - z_2|^{1/2}\). If \(\rho_2 \leq |z_1 - z_2|^{1/2}\) then we have that
\[
\rho_1 = \rho_2 + O(|z_1 - z_2|^{1/2}) = O(|z_1 - z_2|^{1/2}),
\]
by the 1/2-Hölder continuity of the density. Using (150), we readily conclude that
\[
(\eta \rho) \cdot \eta_2 \lesssim |z_1 - z_2|^{1/2} \eta_2 \rho_2 \leq \rho_2 |z_1 - z_2| \eta_1.
\]

By \(\eta_* = \eta_1\) and \(\eta_2 \lesssim \rho_2 \sqrt{\eta_2} \lesssim \rho_2 |z_1 - z_2|^{1/2}\), using \(\rho_2 \gtrsim \sqrt{\eta_2}\) and \(\eta_2 \leq |z_1 - z_2|\). This proves (150) in this case.

In the opposite case \(\rho_2 > |z_1 - z_2|^{1/2}\), using again the Hölder continuity of the density, we have that
\[
\rho_1 = \rho_2 + O(|z_1 - z_2|^{1/2}) \lesssim \rho_2.
\]
Hence we conclude that
\[
(\eta \rho) \cdot \eta_2 \lesssim \eta_* \rho_2 \eta_2 \leq \rho_2 |z_1 - z_2| \eta_1,
\]
where we used \(\eta_* = \eta_1\) and \(\eta_2 \lesssim \eta^* \lesssim |z_1 - z_2|\), proving (150) in this case as well. \(\square\)

Appendix C. Proof of remaining estimates for Proposition 3.4

Proof of the local laws and bound with transposes in Proposition 3.4. Using the bound for \((G_1^i G_2^i A)\) in (43) as an input, and the bound
\[
\langle W G_2^i G_2^i G_1^i A \rangle = O^2 \left( \frac{\sqrt{\rho_1 \rho_2}}{\sqrt{N \eta_1 \eta_2}} \right)
\]
from (61), we prove the bound for \((G_1^i G_2^i A)\) in (51) analogously to the proof of the bound for \((G_1^i G_2^i A)\) above.

Using the local laws for \((G_1^i G_2^i), (G_1^i A G_2^i A)\) in (43) and (44), respectively, and the bound
\[
\langle |W G_2^i G_2^i | \rangle = O^2 \left( \frac{\sqrt{\rho_1 \rho_2}}{\sqrt{N \eta_1 \eta_2}} \right)
\]
from (62) as an input, the proof of the local law for \((G_2^i A G_1^i)\) in (50) follows exactly in the same way as the proof of the local law for \((G_2^i A G_1^i)\) above.

We are now only left with the proof of the local law for \((G_1^i G_2^i)\) in (49). The proof of the local law
\[
\langle G_1^i G_2^i \rangle = \frac{m_1 m_2^i}{(1 - \sigma m_1 m_2)^2} + O_\prec \left( \frac{\rho_1}{\sqrt{N \eta_1 \eta_2}} \right),
\]
for \(\sigma = \pm 1\) requires exactly the same changes to the proof of the local law for \((G_1^i G_2^i)\) that the proof of \((G_1^i G_2^i)\) required to the proof \((G_1^i G_2^i)\). For this reason we omit the details of this proof.

In the remainder we will prove that
\[
\langle G_1^i G_2^i \rangle = \frac{m_1 m_2^i}{(1 - \sigma m_1 m_2)^2} + O_\prec \left( \frac{\rho_1}{\sqrt{N \eta_1 \eta_2}} \right).
\]

The proof of (153) will be divided into three cases: (i) \(\sigma = 1\), (ii) \(\sigma = -1\), (iii) \(|\sigma| < 1\). The case \(\sigma = 1\) is trivial since \(G_1^i = G_1\), hence (153) follows by the local law for \((G_1^i G_2^i)\) in (47). When \(\sigma = -1\) we can write \(W = D + iO\), with \(D\) a diagonal matrix and \(O\) being an skew-symmetric matrix. We now consider the regimes (i) when either \(3z_1 z_2 > 0\) or \(3z_1 z_2 < 0\) and \(|z_1 + z_2| \geq \eta^*\), (ii) \(3z_1 z_2 < 0\) and \(|z_1 + z_2| < \eta^*\). For the case (ii) the local law in (52), together with (76) the bound (82), gives (153). In the regime (i) we consider \(R(z_i) := (iO - z_i)^{-1}\), then using Lemma 3.7 we have
\[
\langle G(z_1)^i G(z_2) \rangle = \langle R(z_1)^i R(z_2) R(z_2) \rangle + O_\prec \left( \frac{1}{N \eta_1 \eta_2 |z_1 + z_2|} \right),
\]
\[
= -\langle R(-z_1) R(z_2) R(z_2) \rangle + O_\prec \left( \frac{1}{N \eta_1 \eta_2 |z_1 + z_2|} \right).
\]
where we used that $R(z_1) = - R(-z_1)$. The proof of the local law
\[- \langle R(z_1) R(z_2) \rangle = - \frac{m(-z_1) m(z_2)}{(1 + m(-z_1) m(z_2))^2} + O_{\prec} \left( \frac{1}{N \eta \eta_2 |z_1 + z_2|} \right) \]
(55)
where we used $m(-z_1) = - m(z_1)$, follows exactly as the proof of the local law for $\langle G_1 G_2 \rangle$ in (79)–(82) and so omitted. Combining (54) and (55) we get
$$\langle G_1^t G_2 \rangle = \frac{m(z_1) m(z_2)}{(1 - m(z_1) m(z_2))^2} + O_{\prec} \left( \frac{1}{N \eta \eta_2 |z_1 + z_2|} \right),$$
which, together with the bound
$$\frac{1}{N \eta \eta_2 |z_1 + z_2|} \lesssim \frac{\rho^*}{L \eta \eta_2}$$
from Lemma 3.8, concludes the proof of the local law for $\langle G_1^t G_2 \rangle$ in (49) for $\sigma = -1$.

In order to conclude the proof of Proposition 3.4 we are now left with the proof of the local law for $\langle G_1^t G_2 \rangle$ in (49) when $|\sigma| < 1$. In this last remaining case, similarly to (81), we use the equation for $\langle G_1^t G_2 \rangle$ and conclude that
\[
\left[ 1 + O_{\prec} \left( \frac{1}{N \eta^*} \right) \right] (1 - \sigma m_1 m_2) \langle G_1^t G_2 \rangle = m_1 \langle G_2 \rangle + \sigma m_1 \langle G_1^t G_2 \rangle \langle G_2 \rangle
\]
\[
- m_1 \langle W G_1^t G_2 \rangle + O_{\prec} \left( \frac{\sqrt{\rho_1 \rho_2}}{\sqrt{L \eta \eta_2}} \right). \tag{56}\]
The bound of the error term in (56) follows by a Schwarz inequality and [5, Lemma 5], similarly to (83) above. Then, using the local laws for $\langle G_2 \rangle$, $\langle G_1^t G_2 \rangle$ in (42) and (43), respectively, we conclude
\[
\left[ 1 + O_{\prec} \left( \frac{1}{N \eta^*} \right) \right] (1 - \sigma m_1 m_2) \langle G_1^t G_2 \rangle = \frac{m_1 m_2'}{(1 - \sigma m_1 m_2)^2} + O_{\prec} \left( \frac{\sqrt{\rho_1 \rho_2}}{\sqrt{L \eta \eta_2}} + \frac{\rho^*}{\sqrt{L \eta \eta_2}} \right)
\]
\[
= \frac{m_1 m_2'}{(1 - \sigma m_1 m_2)^2} + O_{\prec} \left( \frac{\rho^*}{\sqrt{L \eta \eta_2}} \right).
\]
This concludes the proof of the local law for $\langle G_1^t G_2 \rangle$ for the last remaining case $|\sigma| < 1$. \qed

Appendix D. Proof of the refined bounds on renormalised alternating chains: Theorem 3.5 and Lemma 3.6

D.1. Proof of Theorem 3.5. In [5] we proved a general high probability bound [5, Theorem 5] on renormalized alternating chains of resolvents and deterministic matrices. The proof of [5, Theorem 5] proceeded by a delicate power counting of high moments $E(W G_1 B_1 \cdots G_1 B_2)^{2p}$ after iterated cumulant expansions. For this appendix we assume that the reader is familiar with the proof of [5, Theorem 5] and here we only explain the two minor modifications that are necessary for the current setup. The proof of Theorem 3.5 follows directly from the proof of [5, Theorem 5] by replacing each $G_k$ in the formulas [5, Eq. (48)–(50)] by $G_{k,1} \cdots G_{k,n_k}$.

(mod-i) [5, Lemma 5] remains valid if on the lhs. of [5, Eq. (104a)–(104b)] some $G_k$’s are replaced by $G_{k,1} \cdots G_{k,n_k}$ and the rhs. are multiplied by factors of $\min_k \eta_{k,i} \prod_i \eta_{k,i}$ for each replacement.

(mod-2) [5, Lemmata 3–4] hold true under the setting of Theorem 3.5 upon multiplying the rhs. of [5, Eq. (101)–(102b)] by
$$\prod_{k \in [t]} \left( \frac{\min_k \eta_{k,1}}{\prod_i \eta_{k,i}} \right)^{2p}. \tag{57}$$
From (mod-2) the proof of Theorem 3.5 can be concluded exactly as the one of [5, Theorem 5].

Regarding (mod-i) recall that in the proof of [5, Lemma 5] we used three types of estimates on resolvents; (i) the naive isotropic bound $|G_{x,y}| < 1$, (ii) the Ward improvement $(GG^*)_{x,x} < \rho^*/\eta^*$ and
(iii) the norm bound \(|G| \leq 1/\eta_*\). The norm bound is obviously compatible with the replacement by sub-multiplicativity of the norm. For the naive isotropic bound we obtain

\[
|\langle G_1 \cdots G_n \rangle_{xy}| < \min_{i \in \{1,2,3\}} \frac{\eta_i}{\prod_{i \in \{1,2,3\}} \eta_i},
\]

by spectral decomposition. Finally, for the Ward improvement we have

\[
|\langle G_1 \cdots G_n \rangle (G_1 \cdots G_n)^*|_{ax} = \left( \frac{\langle G_1 \cdots G_{n-1} 3 G_n G_{n-1} \cdots G_1 \rangle}{\eta_n} \right) \lesssim \frac{\rho^n \eta_*}{\prod_{i \in \{1,2,3\}} \eta_i} = \frac{\rho^n}{\eta_*} \left( \frac{\eta_*}{\prod_{i \in \{1,2,3\}} \eta_i} \right)^2,
\]

hence also the Ward improvement is compatible with the replacement.

Regarding (mod-2), note that before the cumulant expansions there are 2p copies of each of \(G_{k,1} \cdots G_{k,\alpha_k}\) for \(k = 1, \ldots, l\). However, along the cumulant expansions the initial products of \(G\)'s may be broken up into two or more shorter chains upon differentiation, similarly as a single \(G_{xy}\) becomes \(G_{xu} G_{uy}\). The key additional observation is that the bound (158) is compatible with the cumulant expansion in the sense that the bound cannot increase upon differentiation. For example, "differentiating" the bound

\[
|\langle G_1 G_2 G_3 \rangle_{xy}| \leq \min_{i \in \{1,2,3\}} \frac{\eta_i}{\prod_{i \in \{1,2,3\}} \eta_i}
\]

we get

\[
|\langle G_1 G_2 \Delta^{ab} G_2 G_3 \rangle_{xy}| = |\langle G_1 G_2 \rangle_{xy}| |\langle G_2 G_3 \rangle_{xy}| \lesssim \min_{i \in \{1,2,3\}} \frac{\eta_i}{\prod_{i \in \{1,2,3\}} \eta_i} \left( \min_{i \in \{1,2,3\}} \frac{\eta_i}{\prod_{i \in \{1,2,3\}} \eta_i} \right) \leq \min_{i \in \{1,2,3\}} \frac{\eta_i}{\prod_{i \in \{1,2,3\}} \eta_i}.
\]

Thus the additional factor (157) is an upper bound on the product of additional factors obtained from each application of [15, Lemma 5] in the proofs of [15, Lemmas 3-4].

D.2. Proof of Lemma 3.6. We start with the following bound on \(\langle x, GAWG y \rangle\):

\[
\sqrt{\mathbb{E} \left| \langle x, GAWG y \rangle \right|^2} < \frac{\rho}{\sqrt{N} \eta}, \quad \eta \lesssim 1.
\]

Note that by the bound by (46) it follows that \(\sqrt{\mathbb{E} \left| \langle x, GAWG y \rangle \right|^2} \lesssim \rho^{1/2} \eta^{-1/2}\). In particular, we need an additional \(\sqrt{\rho/(N \eta)}\) gain compared to this bound. To do so we will use (46) as an input to prove a better bound (159) on \(\sqrt{\mathbb{E} \left| \langle x, GAWG y \rangle \right|^2}\).

Proof of (159). From now on, without loss of generality, we assume that \(A = A^*\). To make the presentation cleaner we assume that \(w_2 = 1 + \sigma\), the general case \(w_2 \neq 1 + \sigma\) is completely analogous and so omitted, since the difference compared to the case presented here is only in the case \(b = a\) in the summations that is of lower order.

Using cumulant expansion, and the notation \(\alpha := (\alpha_1, \ldots, \alpha_k)\) we start computing

\[
\mathbb{E} \langle x, GAWG y \rangle \langle y, G^* WAG^* x \rangle = \frac{1}{N} \mathbb{E} \sum_{ab} \langle x, GAD^{ab} G y \rangle \left( \langle y, G^* \Delta^{ba} AG^* x \rangle - \langle y, G^* \Delta^{ba} G^* WAG^* x \rangle \right) + \ldots
\]

\[
+ \frac{\rho}{N} \mathbb{E} \sum_{ab} \langle x, GAD^{ab} G y \rangle \left( \langle y, G* \Delta^{ab} AG^* x \rangle - \langle y, G^* \Delta^{ab} G^* WAG^* x \rangle \right) + \ldots
\]

\[
+ \sum_{k \geq 2} \sum_{ab} \sum_{\alpha \in \{ab, ba\}^k} \frac{k(ab, \alpha)}{k!} \mathbb{E} \partial_\alpha \left[ \langle x, GAD^{ab} G y \rangle \langle y, G^* WAG^* x \rangle \right].
\]

The dots in the second and third line of (160) denote the fact that we considered only the case when the derivative hits the first \(G^*\) in \(\langle y, G^* WAG^* x \rangle\). The case when the derivative hits the second \(G^*\) is completely analogous and so omitted. We will often use this notation in the remainder of the proof to denote that we consider only some representative terms and all the others are estimated analogously.

In order to perform cumulant expansion in the remaining underlined term in the last line of (160) we consider the case when either one of the \(\alpha\)-derivative hits the \(W\) or all the derivatives hit a \(G^*\). Without
loss of generality we denote this derivative by $\partial_{\alpha_1}$ and we use the notation $\alpha_2 := (\alpha_2, \ldots, \alpha_k)$. Then, performing cumulant expansion in the last line of \eqref{eq:160}, we get

$$
\begin{align*}
E(x, GA^*W G^* y) & = \frac{1}{N} \mathbb{E}(x, G A^* y) + \frac{1}{N^2} \mathbb{E}(x, G A^* G^* y) + \frac{1}{N^3} \mathbb{E}(x, G A^* G^* G^* y) + \ldots \\
+ \frac{1}{N} \sum_{abcd} \mathbb{E}(x, G A^* G^* y) + \frac{1}{N^2} \sum_{abcd} \mathbb{E}(x, G A^* G^* G^* y) + \frac{1}{N^3} \sum_{abcd} \mathbb{E}(x, G A^* G^* G^* G^* y) + \ldots \\
+ \frac{1}{N^2} \sum_{abcd} \mathbb{E}(x, G A^* G^* G^* y) + \frac{1}{N^3} \sum_{abcd} \mathbb{E}(x, G A^* G^* G^* G^* y) + \ldots \\
+ \frac{1}{N^3} \sum_{abcd} \mathbb{E}(x, G A^* G^* G^* G^* y) + \frac{1}{N^4} \sum_{abcd} \mathbb{E}(x, G A^* G^* G^* G^* G^* y) + \ldots \\
+ \ldots \\
+ \ldots \\
+ \ldots
\end{align*}
$$

Here we omitted the third or higher order terms coming from the cumulant expansions in the third and fourth line of \eqref{eq:160}, since they are estimated exactly as the ones in the sixth and seventh line of \eqref{eq:160}.

We now first estimate the last three lines of \eqref{eq:160}. We rewrite the derivative in the third to last line together with the $a$-summation (which can be performed as a matrix product) as

$$
\frac{1}{N} \sum_{abcd} \sum_{k \geq 2} \sum_{ab \in \{ab, ba\}^k, \alpha_1 \in \{ab, ba\}} \frac{\kappa(ab, \alpha_1, \alpha_2)\kappa(cd, \beta)}{k!} \partial^a E \left[ \langle x, G A^* G^* y \rangle \langle y, G^* A^* G^* y \rangle \right] + \ldots
$$

and notice that irrespective of $\beta$ and the allocation of derivatives we have one factor of $G A^*$, one factor of $G$ evaluated in one of $by, bc, bd$, one factor of $G^*$ evaluated in one of $yb, cb, db$ and one factor of $AG^*$ or $G^*$ evaluated in one of $dx, ca$. Thus, in any case three Ward estimates can be performed (two in the $b$-index and one in the $c$- or $d$-index) and together with the estimate $\| (G A^* G^*)_{xy} \| \leq \sqrt{\rho/\eta}$ and the scaling $|\kappa(cd, \beta)| \lesssim N^{-(k+1)/2}$ we obtain a bound of at most

$$
N^{1-(k+3)/2} \sqrt{\frac{\rho}{\eta}} \left( \frac{\rho}{N \eta} \right)^{3/2} = \frac{\rho^2}{N^{1/2} \eta^2} \leq \frac{\rho^2}{N \eta^2}.
$$

For the second to last and last line of \eqref{eq:160} the argument is considerably simpler since a simple counting of available Ward estimates suffices. For the $k$-th term of the penultimate line of \eqref{eq:160} the naive size is given by $N^{2-(k+1)/2}$ (i.e. using $|G_{xy}| \leq 1$ for all factors), while for each derivative allocation at least four resolvent entries can be estimated via the Ward identity, hence gaining a factor of $\rho^2 / (N^2 \eta^2)$, yielding $\rho^2 N^{2-(k+1)/2} \eta^{-2}$. For the $k$-th term of the last line of \eqref{eq:160} the naive size is given by $N^{3-(k+1)/2}$, while for each derivative allocation at least five resolvent entries can be estimated via the Ward identity, hence gaining a factor of $\rho^5 / (N \eta)^{5/2}$, yielding $N^{3-(k+1)/2} \rho^5 \eta^{-5/2}$. Thus, we finally conclude that the last three lines of \eqref{eq:160} are bounded by $\rho^2 N^{-1} \eta^{-2}$.
In order to conclude the proof of (159) we are left only with the second order terms, i.e. we are left with
\[
\begin{align*}
E(x, GAWGy) (y, G^* W A^G x) &= \frac{E(GG^* y y (G A^2 G^*)_{xx}}{N} + \frac{\sigma}{N} E(G^* A^2 G^*)_{yy} (G^* AG^*)_{yx} + \frac{E(GA^2 G^*)_{xy} (G^* GAG^*)_{yx}}{N^2} \\
&+ \frac{1}{N} E(G^* G^*)_{yy} (G^* G^*)_{xx} + \frac{\sigma}{N^2} E(G(G^*)^4 A^2 G^* AG^*)_{xx} (G^* G)_{yy} \\
&+ \frac{\sigma^2}{N} E(G(G^*)^4 y y (G^* AG^*)_{yx} + \frac{\sigma^2}{N^2} E(GAG^* G^* G^* G^*)_{xx} (G^* AG^*)_{yx} \\
&+ \frac{\sigma^2}{N^2} E(G^* A^2 G^* G^*)_{xy} (G^* G^* G^*)_{yx} + O \left( \frac{\rho^2 N^2}{N^2 \eta^2} \right).
\end{align*}
\]
for any \( \epsilon > 0 \). In order to conclude (159) we need to show that all the terms in (162) are bounded by \( \rho^2 / (N \eta^2) \). For this purpose we need an additional bound which improves the bound in [5, Lemma \( \xi \)]. We present this bound in the following lemma, which will be proven at the end of this section.

**Lemma D.1.** We have the bound
\[
|G^* A^2 G^* AG^*| < \frac{N \rho}{\eta}.
\]
(163)

Note that by a Schwarz inequality (after the first \( G \)) and the bound in [5, Lemma \( \xi \)] with \( \Lambda_+ \lesssim 1 \), we conclude that
\[
|\langle G(G^*)^2 A^2 AG^* \rangle_{xx}| < \frac{N \rho}{\eta}.
\]
(164)

We also have that
\[
|\langle GAG^* G^* G^* \rangle_{xy}| \leq \frac{1}{\eta} \langle GAG^* \rangle_{xx}^{1/2} \langle G^* AG^* \rangle_{yy}^{1/2} < \frac{\rho \sqrt{N}}{\eta^2},
\]
(165)

that \( \langle GCG^* \rangle \leq \|C\| \langle G \rangle / \eta \), for any matrix \( C = G^* \), and that \( \|\langle G \rangle\| \leq 1/\eta \). Inserting the bounds (163), (164), and (165) into (162) we conclude that
\[
E(x, GAWGy) (y, G^* W A^G x) = O \left( \frac{N^2 \rho^2}{N^2 \eta^2} \right),
\]
for any \( \epsilon > 0 \). This concludes the proof of (159).

\[
\square
\]

The proof of the bounds in (60)–(62) is very similar to the one presented above, hence we will present the estimate only of some representative terms. In particular, the higher order cumulants can be estimated in the bound on the last three lines of (162) above and we therefore here only consider the more critical second order cumulant terms. Also note that the bounds (60)–(62) follow directly from [5, Theorem \( \xi \)] in case when \( \rho_1, \rho_2 \) and \( \eta_1, \eta_2 \) are comparable, hence the purpose of the variance calculation is just to obtain the individual dependences stated in (60)–(62). We now present the proof of the underlined terms in (60)–(62) containing transposes, the proof for the underlined terms without transposes is completely analogous and so omitted.

**Proof of (60).** We compute the second moment performing cumulant expansion exactly as in (160). Here we write only some representative second order terms:
\[
E[|WG_1 G_2^* A|^2] = E \left[ \frac{1}{N} (G_1 G_2^* A^2 (G_2^*)^* G_1^*) + \frac{1}{N^2} (G_1^* G_2^* G_1^* G_2^* A^2) + \frac{1}{N^2} (G_1 G_2^* A (G_2^*)^* G_1^*) G_2^* A^2 \right] \\
+ \ldots
\]
(167)
Using that $\|A^2\| \lesssim 1$ and Ward identity $\eta_1 G_1 G_1^* = \exists G_1$, we conclude that the first term in (167) is bounded by
\[
\frac{1}{N^2} \langle G_1 G_2^* A^2 (G_2^*)^* G_1^* \rangle \leq \frac{1}{N^2 \eta_1 \eta_2} \langle 3 G_1 \exists G_2^* \rangle \times \rho_* \frac{\rho_*}{N^2 \eta_1 \eta_2},
\]
where in the last inequality we used $|\langle \exists G_1 \rangle| \prec \rho_*$. For the second term in (167) we use Schwarz inequality to separate resolvents with and without transpose, then we use Ward identity, to obtain
\[
\frac{1}{N^2} \langle G_1^* G_1^* \rangle \langle G_1 G_2^* A (G_1^*)^* G_2^* A^* \rangle \times \frac{\rho_1}{N^2 \eta_1} \langle G_1 A^* G_2^* A^* G_1^* \rangle \times \frac{\rho_1^2 \rho_2}{N^2 \eta_1 \eta_2}.
\]
where we used that by [5, Lemma 5] it holds $\langle \exists G_1 A^* \exists G_2 A^* \rangle \prec \rho_1 \rho_2$. For the third term in (167) we bound
\[
\frac{1}{N^2} \langle \langle G_1 G_2^* A G_1^* \rangle^2 \rangle \leq \frac{1}{N^2} \langle G_1 G_1^* \rangle \langle G_1 A (G_2^*)^* G_2^* A G_1^* \rangle \times \frac{\rho_1 \rho_2}{N^2 \eta_1 \eta_2},
\]
again by [5, Lemma 5]. We now bound the fourth term in (167):
\[
\frac{1}{N^1} \langle G_1 A (G_2^*)^* G_1^* A^* (G_2^*)^* G_1^* \rangle \leq \frac{1}{N^1 \eta_1} \langle G_1 A (G_2^*)^* A^* (G_2^*)^* G_1^* \rangle \langle G_1 A (G_2^*)^* G_1^* A^* (G_2^*)^* G_1^* \rangle \times \frac{\rho_1 \rho_2}{N^1 \eta_1 \eta_2}.
\]
This concludes the bound for the second order terms in the cumulant expansion.

\[\square\]

Proof of (61). The proof of (61) is analogous to the proof of (60), we just pick up an additional $1/\eta_1$ factor due to the one more $G_1$ term. We only show the bound for a few second order terms:
\[
\mathbb{E}|\langle W G_1 G_2^* \rangle|^2 = \mathbb{E} \frac{1}{N^2} \langle G_1 G_2^* A^2 (G_2^*)^* G_1^* \rangle
\]
\[
+ \mathbb{E} \frac{1}{N^2} \langle G_1 G_2^* A (G_1^*)^* G_2^* A^* \rangle \langle G_1 G_2^* A (G_1^*)^* G_2^* A^* \rangle + \ldots
\]
(69)

For the first term in (69) we have that
\[
\frac{1}{N^2} \langle G_1 G_2^* A^2 (G_2^*)^* G_1^* \rangle \leq \frac{1}{N^2 \eta_1 \eta_2} \langle 3 G_1 \exists G_2^* \rangle \times \rho_* \frac{\rho_*}{N^2 \eta_1 \eta_2},
\]
For the second term in (69) we bound
\[
\frac{1}{N^2} \langle G_1 G_2^* A (G_1^*)^* G_2^* A^* \rangle \times \frac{\rho_1}{N^2 \eta_1 \eta_2} \langle G_1 A^* G_2^* A^* G_1^* \rangle \times \frac{\rho_1^2 \rho_2}{N^2 \eta_1 \eta_2}.
\]
Similarly to (68), for the third term in (69) we get
\[
\frac{1}{N^2} \langle G_1 A (G_2^*)^* A^* (G_2^*)^* G_1^* \rangle \times \frac{\rho_1 \rho_2}{N^2 \eta_1 \eta_2}.
\]
This concludes the bound for the second order terms in the cumulant expansion.

\[\square\]

Proof of (62). We consider a few representative second order terms from the cumulant expansion
\[
\mathbb{E}|\langle W G_1 G_2^* \rangle|^2 = \mathbb{E} \frac{1}{N^2} \langle G_1 G_2^* A G_1^* A^* G_2^* A (G_2^*)^* G_1^* \rangle + \mathbb{E} \frac{1}{N^2} \langle G_1 G_2^* A G_1^* A^* G_2^* A (G_2^*)^* G_1^* \rangle^2
\]
\[
+ \mathbb{E} \frac{1}{N^2} \langle G_1 G_2^* A G_1^* A^* G_2^* A (G_2^*)^* G_1^* \rangle \times \frac{\rho_1}{N^2 \eta_1 \eta_2} \langle G_1 A^* G_2^* A^* G_1^* \rangle \times \frac{\rho_1^2 \rho_2}{N^2 \eta_1 \eta_2} + \ldots
\]
(70)
The first term in (170) is bounded by
\[
\frac{1}{N^2} \langle G_1 G_2 A G_1 A^2 G_1^* A (G_2^*)^* G_1^* \rangle \leq \frac{1}{N^{2\eta_1^2}} \langle 3G_1 G_2 A 3G_1 A (G_2^*)^* \rangle \\
\leq \frac{1}{N^{2\eta_1^2}} \sqrt{\langle 3G_1 3G_1 \rangle \langle G_2^* A 3G_1 A (G_2^*)^* G_2^* A 3G_1 A (G_2^*)^* \rangle} \\
\leq \frac{\rho_{1/2} \rho_{2}}{N^{3/2} \eta_1^2 \eta_2}.
\]

By a Schwarz inequality and [15, Lemma 5] we readily conclude that the second term in (170) is bounded by
\[
\frac{1}{N^2} |\langle G_1 G_2 A G_1 A^2 G_1^* \rangle|^2 < \frac{\rho_{1} \rho_{2}}{N \eta_1^2 \eta_2}.
\]

We now bound the third term in (170):
\[
\frac{1}{N^2} \langle G_1 G_2 A G_1 A (G_1^*)^* G_2^* A (G_1^*)^* A^4 \rangle \leq \frac{\rho_{1} \rho_{2}}{N \eta_1^2 \eta_2} (3G_1^* B G_2^* A 3G_2 A G_1 A)
\]
\[
\leq \frac{\rho_{1} \rho_{2}}{N \eta_1^2 \eta_2},
\]
where we used a Schwarz inequality and [15, Lemma 5].

We conclude this section with the proof of Lemma D.1.

**Proof of Lemma D.1.** Similarly to (77), writing the equation for \( G^* A^4 G^* A^* \) we conclude that
\[
\begin{aligned}
&\left[1 + \mathcal{O}_\kappa \left( \frac{1}{N \eta} \right) \right] \langle G^* A^4 G^* A^* \rangle_{xy} \\
&= \bar{n}(A^4 G^* A^*)_{xy} - \bar{m}(WG^* A^4 G^* A^*)_{xy} + \sigma_{1} \bar{m}(G^* A^4 G^*)_{xy} \\
&\quad + \bar{m}(G^*)_{xy} (G^* A^4 G^* A^*)_{xy} + \frac{\bar{m}_x}{N} (\text{diag}(G^* A^4 G^*)_{xy} + \frac{\bar{m}_x}{N} (\text{diag}(G^* A^4 G^* A^*)_{xy} \\
&\quad + \frac{\bar{m}_w}{N} (\text{diag}(G^* A^4 G^* A^*)_{xy} + \frac{\bar{m}_r}{N} (G^* A^4 G^* A^*)_{xy} \\
&\quad + \frac{\bar{m}_w}{N} (\text{diag}(G^* A^4 G^* A^*)_{xy} = \mathcal{O}_\kappa \left( \sqrt{\frac{N \rho}{\eta}} \right),
\end{aligned}
\]

where we used a Schwarz inequality and [15, Lemma 5] (similarly to (83)) to bound all the terms with a pre-factor \( N^{-1} \), and that
\[
\langle W G^* A^4 G^* A^* \rangle_{xy} \leq \sqrt{\frac{N \rho}{\eta}}, \quad \langle G^* A^4 G^* A^* \rangle \sim \frac{1}{\sqrt{\eta}}, \quad \langle G^* A^4 G^* A^* \rangle \sim \frac{\rho}{\eta},
\]
with the first bound by (63), and the last two bounds from [15, Lemma 5] with \( \Lambda_+ \lesssim 1 \), and that
\[
\langle G A G^* \rangle_{xy} \sim \sqrt{\rho},
\]
by (46).
Appendix E. Calculations for the functional CLT for smooth test functions

Starting from (140) and using the explicit formulas in Theorem 4.1 in this section we complete the proof of Theorem 2.4 by computing the expectations and variances of the limiting Gaussian processes explicitly. In the following we will often use that by Stokes Theorem it follows

$$\int_\mathbb{R} \int_{\eta_1}^{\eta_2} \partial_\eta \psi(x + i\eta) h(x + i\eta) \, dx \, d\eta = \frac{1}{2\pi} \int_\mathbb{R} \psi(x) h(x) \, dx$$ (171)

for any $\eta_1 \in [0, \eta_0]$, and for any $\psi, h \in H^1$ such that $\partial_\eta h = 0$ on the domain of integration and for $\psi$ vanishing at the left, right, and top boundary of the domain of integration.

We first compute the order $N^{-1}$ correction to the expectation in Section E.1, then in Section E.2 we prove an approximate Wick Theorem for $L_N(f, I)$, $L_N(f, A_d)$, $L_N(f, A_{od})$, and, finally, we explicitly compute their variance.

E.1. Computation of the expectation. By (93), using the notation $m = m_{sc}$, for any $z = x + i\eta \in \mathbb{C} \setminus \mathbb{R}$ with $\eta \gg N^{-1}$, and for any $\epsilon > 0$, it follows that

$$\langle E G(z) \rangle = m + \frac{\sigma m^2}{N m (1 - \sigma m^2)} + \frac{\kappa_4}{N} \frac{m'}{m} + \frac{\bar{w}_2 m'}{N} + O\left(\frac{N^\epsilon}{(N\eta)^{3/2}}\right),$$ (172)

$$\langle E G(z) \hat{A} \rangle = O\left(\frac{N^\epsilon \rho_{sc}(z)^{1/2}}{N^{3/2} \eta}\right).$$

In the following computations we will often omit the $N^\epsilon$ factor in the error terms. By the second line of (172), together with (34), we readily conclude that

$$N^{1+\alpha/2} \int_\mathbb{R} \int_{\eta_0}^{\eta_0} \partial_\eta f_G(z) \, E G(x + i\eta) \, dx \, d\eta \lesssim N^{-(1-\alpha)/2} \log N.$$ (173)

This concludes the bound for the expectation of $L_N(f, \hat{A})$.

Next we proceed with the computation of $E L_N(f, I)$. Using the first equality in (172), and the bound (34) to estimate the error term, we get that

$$\frac{2N}{\pi} \int_\mathbb{R} \int_{\eta_0}^{\eta_0} \partial_\eta f_G(z) \left[ m + \frac{\sigma m m'}{N (1 - \sigma m^2)} + \frac{\kappa_4}{4N} \partial_x (m^4) + \frac{\bar{w}_2 m'}{N} \right] \, d\eta \, dx + O\left(N^{-\alpha/2}\right).$$ (174)

Then, by (179), it is easy to see that

$$\frac{2N}{\pi} \int_\mathbb{R} \int_{\eta_0}^{\eta_0} \partial_\eta f_G(z) \, m \, dy \, dx = N \int_{-2}^{2} f(x) \rho_{sc}(x) \, dx + O\left(N^{1+\alpha/2}\theta_0^2\right).$$

In the following computations we will often use that the regime $\eta \in [\eta_-, \eta_0]$, with $\eta_0 := N^{-100}$, is added back to the integration in (174) at the price of a negligible error much smaller than $N^{-(1-\alpha)/2}$.

In the following by $\log z$ we denote the complex logarithm on $\mathbb{C} \setminus \mathbb{R}^+$. For the second term in (174), by (177), using that for analytic functions $h$ it holds $\partial_\eta h = -i\partial_x h$, and that

$$\partial_x \log(1 - \sigma m^2) = -\frac{2\sigma m m'}{1 - \sigma m^2},$$
we compute

\[
\frac{1}{\pi} \int_R \int_{\eta \geq 0} \partial_x f_C(z) \partial_\eta \log(1 - \sigma m^2) \, dx \, d\eta
\]

\[
= \frac{1}{\pi} \int_R \int_{\eta \geq 0} \partial_x \eta f_C(z) \log(1 - \sigma m^2) \, dx \, d\eta + O(\eta \log \eta)
\]

\[
= \frac{1}{2\pi} \Im \lim_{\epsilon \to 0+} \left( \int_{-\epsilon}^{\epsilon} + \int_{\epsilon}^{2-\epsilon} \right) f'(x) \log \left[ 1 - \sigma m^2(x + \im \eta) \right] \, dx + O(\sqrt{\eta})
\]

\[
= \frac{1}{2\pi} \Im \left[ f(2 - \epsilon) \log \left( 4 - \sigma(2 - \epsilon)^2 + \im \sigma(2 - \epsilon)\sqrt{4 - (2 - \epsilon)^2} \right)
\]

\[
- f(-2 + \epsilon) \log \left( 4 - \sigma(2 - \epsilon)^2 - \im \sigma(2 - \epsilon)\sqrt{4 - (2 - \epsilon)^2} \right)
\]

\[
- \frac{1}{2\pi} \Im \lim_{\epsilon \to 0+} \left[ f(2 - \epsilon) \log(2(1 + \sigma) - \sigma \epsilon^2 + \im \sigma \epsilon \sqrt{4 - \epsilon^2})
\]

\[
- f(-\epsilon) \log(2(1 + \sigma) - \sigma \epsilon^2 - \im \sigma \epsilon \sqrt{4 - \epsilon^2})
\]

\[
+ \frac{1}{\pi} \int_R f(x) \frac{\sigma m(x + \im \eta)m'(x + \im \eta)}{1 - \sigma m(x + \im \eta)^2} \, dx + O(\sqrt{\eta})
\]

\[
= \frac{1}{2\pi} \Im \left[ f(2 - \epsilon) \arctan \left( \frac{\sigma (2 - \epsilon) \sqrt{4 - (2 - \epsilon)^2}}{4 - \sigma (2 - \epsilon)^2} \right)
\]

\[
- f(-2 + \epsilon) \arctan \left( -\frac{\sigma (2 - \epsilon) \sqrt{4 - (2 - \epsilon)^2}}{4 - \sigma (2 - \epsilon)^2} \right)
\]

\[
- \frac{1}{2\pi} \Im \left[ f(\epsilon) \arctan \left( \frac{\sigma \epsilon \sqrt{4 - \epsilon^2}}{2(1 + \sigma) - \sigma \epsilon^2} \right) - f(-\epsilon) \arctan \left( -\frac{\sigma \epsilon \sqrt{4 - \epsilon^2}}{2(1 + \sigma) - \sigma \epsilon^2} \right) \right]
\]

\[
- \frac{1}{2\pi} \int_{-2}^{2} f(x) \frac{\sigma [2(1 + \sigma) - x^2]}{[1 + \sigma]^2 - \sigma x^2 \sqrt{4 - x^2}} \, dx
\]

\[
= 1(\sigma = 1) \frac{f(2) + f(-2)}{4} + 1(\sigma = -1) \frac{f(0)}{2} - \frac{1}{2\pi} \int_{-2}^{2} f(x) \sigma [2(1 + \sigma) - x^2] \, dx + O(\sqrt{\eta}).
\]

To go from the second to the third line we used that \( \Im \log[1 - \sigma m(x + \im \eta)] = O(\sqrt{\eta}) \) for \( x \in [-2, 2]^c \).

For the third term in (74), using (77), we have

\[
\frac{\kappa_4}{2\pi} \int_R \int_{\eta_0} \partial_x f_C(z) \partial_x (m^2) \, d\eta \, dx = \frac{\kappa_4}{2\pi} \Im \int_R f(x) \frac{x^4 - 4x^2 + 2}{\sqrt{4 - x^2}} \, dx + O(\sqrt{\eta}).
\]

Finally, for the last term in (74), using (77) again, we conclude

\[
\frac{\mu_2}{2\pi} \Im \int_R f(x) (m^2)' \, dx = -\frac{\mu_2}{2\pi} \int_{-2}^{2} f(x) \frac{2 - x^2}{\sqrt{4 - x^2}} \, dx.
\]

This concludes the computation of \( \text{E}L_N(f, I) \), and so, together with the bound (73), it concludes the computation of the expectation of \( L_N(f, A) = \langle A \rangle L_N(f, I) + L_N(f, \tilde{A}) \).

### E.2. Wick Theorem and computation of the variance

Now we proceed with the computation of moments of products of \( L_N(f, I), L_N(f, \tilde{A}_u) \) and \( L_N(f, A_{od}) \).
By (91)–(92) we have that
\begin{equation}
E\left(\prod_{i \in [p]} (G(z_i)A_{i,od}) \right) \left(\prod_{i \in (p,q)} (G(z_i)\hat{A}_{i,d}) \right) \left(\prod_{i \in (q,r)} ([G(z_i) - E(G(z_i))]) \right) \nonumber
\end{equation}
\begin{equation}
= \frac{1}{\mathcal{N}} \sum_{P \in \text{Pair}([p]), (i,j) \in P} \prod_{(i,j) \in P} \left[ (A_{i,od}A_{j,od}) \frac{m_i^2 m_j^2}{1 - m_i m_j} + (A_{i,od}A_{j,od}^t) \frac{\sigma m_i^2 m_j^2}{1 - \sigma m_i m_j} \right] \nonumber
\end{equation}
\begin{equation}
\times \prod_{(i,j) \in Q} \langle \hat{A}_{i,d} \hat{A}_{j,d} \rangle \left( \frac{m_i^2 m_j^2}{1 - m_i m_j} + \frac{\sigma m_i^2 m_j^2}{1 - \sigma m_i m_j} + \tilde{w}_2 m_i^2 m_j^2 + \kappa_4 m_i^3 m_j^3 \right) \nonumber
\end{equation}
\begin{equation}
\times \prod_{(i,j) \in R} \left( \frac{m_i^2 m_j^2}{1 - m_i m_j} + \frac{\sigma m_i^2 m_j^2}{1 - \sigma m_i m_j} + \tilde{w}_2 m_i^3 m_j^2 + \frac{\kappa_4}{2} (m_i^2) (m_j^2) \right) \nonumber
\end{equation}
\begin{equation}
+ \mathcal{O}\left( \frac{N^4 \Psi \sqrt{N \eta}}{N \eta} \right), \tag{76}
\end{equation}
where \( \eta_* := \min\{||\eta_i| : i \in [k]\} \), and
\begin{equation}
\Psi := \prod_{i \in [q]} \frac{\rho(z_i)^{1/2}}{N \eta_i^{1/2}} \prod_{i \in [q,r]} \frac{1}{N \eta_i}. \tag{77}
\end{equation}

Here we considered the case when \( p, q \) and \( r \) are even, the case when one of them is odd the leading term is zero. Using (76), together with \( |\eta_0 f_c| \leq N^N |f''| + N^N |(f) + |f'|| \) for \( \eta \in \Gamma_{\eta_0, \eta_*} := [\eta_0, \eta_*] \), by (134), it readily follows that
\begin{equation}
E\left(\prod_{i \in [p]} N^{n_i/2} L_N(f^{(i)}, A_{i,od}) \right) \left(\prod_{i \in (p,q)} N^{n_i/2} L_N(f^{(i)}, \hat{A}_{i,d}) \right) \left(\prod_{i \in (q,r)} L_N(f^{(i)}, I) \right) \nonumber
\end{equation}
\begin{equation}
= \sum_{P \in \text{Pair}([p]), (i,j) \in P} \prod_{(i,j) \in P} \left[ \frac{N^N}{\pi^2} \int_{R} \int_{\Gamma_{\eta_0, \eta_*}} \partial f_{C}^{(i)}(z_i) \partial f_{C}^{(j)}(z_j) \right] \nonumber
\end{equation}
\begin{equation}
\times \left( \langle A_{i,od}A_{j,od} \rangle \frac{m_i^2 m_j^2}{1 - m_i m_j} + \langle A_{i,od}A_{j,od}^t \rangle \frac{\sigma m_i^2 m_j^2}{1 - \sigma m_i m_j} \right) \nonumber
\end{equation}
\begin{equation}
\times \prod_{(i,j) \in Q} \int_{R} \int_{\Gamma_{\eta_0, \eta_*}} \partial f_{C}^{(i)}(z_i) \partial f_{C}^{(j)}(z_j) \nonumber
\end{equation}
\begin{equation}
\times \left( \frac{m_i^2 m_j^2}{1 - m_i m_j} + \frac{\sigma m_i^2 m_j^2}{1 - \sigma m_i m_j} + \kappa_4 m_i^3 m_j^3 + \tilde{w}_2 m_i^2 m_j^2 \right) \nonumber
\end{equation}
\begin{equation}
\times \prod_{(i,j) \in R} \left( \frac{m_i^2 m_j^2}{1 - m_i m_j} + \frac{\sigma m_i^2 m_j^2}{1 - \sigma m_i m_j} + \tilde{w}_2 m_i^3 m_j^2 + \frac{\kappa_4}{2} (m_i^2) (m_j^2) \right) \nonumber
\end{equation}
\begin{equation}
+ \mathcal{O}\left( \frac{N^4 \Psi \sqrt{N \eta}}{N \eta} \right). \tag{78}
\end{equation}
for any \( \epsilon > 0 \). The equality in (78) concludes the proof of the Wick pairing. In the following sections we explicitly compute the integrals in the rhs. of (78). In Section 8.2.1 we compute the variance of \( L_N(f, I) \), then in Section 8.2.2 we compute the variance of \( L_N(f, A_{od}) \) and \( L_N(f, \hat{A}_{d}) \).
E.2.1. Computation of the variance of $L_N(f, I)$. Adding the regime $\eta \in [\eta_\varepsilon, \eta_0]$, with $\eta_\varepsilon := N^{-100}$, at the price of a negligible error much smaller $N^{-(1-\alpha)/2}$, we start with

$$
\frac{1}{\pi^2} \int_{\mathbb{R}} \int_{|\eta| \geq \eta_\varepsilon} \int_{|\eta'| \geq \eta_\varepsilon} \partial_\eta f c(z_1) \partial_\eta g c(z_2) \left[ \frac{m_1^2 m_2^2}{(1 - m_1 m_2)^2} \left( 1 - \sigma m_1 m_2 \right) \right]^{\frac{1}{2}} + \frac{\kappa_4}{2} \left( m_1^2 \right) \left( m_2^2 \right) + \hat{\omega}_2 m_1^2 m_2^2.
$$

By direct computations, using (77), the last term in (79) is given by

$$
\frac{4\hat{\omega}_2}{\pi^2} \left( \int_{\mathbb{R}} \int_{|\eta| \geq \eta_\varepsilon} \partial_\eta f c(z) m' \, d\eta \, dx \right) \left( \int_{\mathbb{R}} \int_{|\eta| \geq \eta_\varepsilon} \partial_\eta g c(z) m' \, d\eta \, dx \right) = \frac{\hat{\omega}_2}{4\pi^2} \left( \int_{-\infty}^{\infty} f(x) \frac{2 - x^2}{\sqrt{4 - x^2}} \, dx \right) \left( \int_{-\infty}^{\infty} g(x) \frac{2 - x^2}{\sqrt{4 - x^2}} \, dx \right).
$$

Then, using that

$$
\int_{\mathbb{R}} \int_{|\eta| \geq \eta_\varepsilon} \partial_\eta f c \partial_\eta z (m^2) \, d\eta \, dx = 2 \int_{\mathbb{R}} \int_{|\eta| \geq \eta_\varepsilon} \partial_\eta f c \partial_\eta z (m^2) \, d\eta \, dx
$$

by (77), we conclude that the $\kappa_4$ coefficient in (79) is given by

$$
\frac{\kappa_4}{2\pi^2} \left( \int_{-\infty}^{\infty} f(x) \frac{2 - x^2}{\sqrt{4 - x^2}} \, dx \right) \left( \int_{-\infty}^{\infty} g(x) \frac{2 - x^2}{\sqrt{4 - x^2}} \, dx \right).
$$

Next, using that

$$
\partial_1 \partial_2 \log(1 - \sigma m_1 m_2) = -\frac{\sigma m_1 m_2}{(1 - \sigma m_1 m_2)^2},
$$

with $z = x + i\eta$, $w = y + i\eta'$, and $m_1 = m_{sc}(z)$, $m_2 = m_{sc}(w)$, for $\sigma = 1$ we compute

$$
\frac{1}{\pi^2} \int_{\mathbb{R}} \int_{|\eta| \geq \eta_\varepsilon} \int_{|\eta'| \geq \eta_\varepsilon} \partial_\eta f c(z) \partial_\eta g c(w) \frac{m_1^2 m_2^2}{(1 - m_1 m_2)^2} + \frac{\kappa_4}{2} \left( m_1^2 \right) \left( m_2^2 \right) + \hat{\omega}_2 m_1^2 m_2^2.
$$

where to go from the third to the fourth line we used (77) and that $(\partial_\eta f c)(x \pm i\eta) = i f'(x)$, since $\eta_\varepsilon \ll N^{-\alpha}$.

Note that

$$
\Re \left[ \log(1 - m_1 m_2) - \log(1 - m_1 m_2) \right] = \mathcal{O}(\sqrt{\eta_\varepsilon}),
$$

$$
\partial_\eta \Re \left[ \log(1 - m_1 m_2) - \log(1 - m_1 m_2) \right] = \mathcal{O}(\sqrt{\eta_\varepsilon}),
$$

(84)
if either $x \in [-2, 2]^c$ or $y \in [-2, 2]^c$. Next, using (88), we compute
\[
\int_{\mathbb{R}} \int_{\mathbb{R}} \nabla_x f(x) \nabla_y g(y) \Re \left[ \log(1 - \sigma m_1 m_2) - \log(1 - m_1 m_2^2) \right]
= - \int_{\mathbb{R}} \int_{\mathbb{R}} \nabla_x f(x) \nabla_y g(y) \Re \left[ \log(1 - \sigma m_1 m_2) - \log(1 - m_1 m_2^2) \right] + O(\sqrt{\eta}).
\]
(88)

Then, using the same computations as in (83) but integrating with respect to the $x$-variable, we conclude
\[
\int_{\mathbb{R}} \int_{\mathbb{R}} \nabla_x f(x) \nabla_y g(y) \Re \left[ \log(1 - \sigma m_1 m_2) - \log(1 - m_1 m_2^2) \right]
= - \frac{1}{2} \int_{\mathbb{R}} \left[ (f(y) - f(x))(g(y) - g(x)) \Re \left\{ \frac{m_1' m_2'}{(1 - \sigma m_1 m_2) + (1 - m_1 m_2^2)} \right\} \right] \Re \left\{ \frac{m_1' m_2'}{(1 - \sigma m_1 m_2) + (1 - m_1 m_2^2)} \right\} dy + O(\sqrt{\eta}).
\]
(86)

Combining (83), and (88)–(86), we finally conclude that the first term the last line of (78) is given by
\[
\frac{1}{4 \pi^2} \int_{\mathbb{R}^2} \left( f(y) - f(x) \right) \left( g(y) - g(x) \right) \Re \left\{ \frac{m_1' m_2'}{(1 - \sigma m_1 m_2) + (1 - m_1 m_2^2)} \right\} dy dx + O(\sqrt{\eta}).
\]
(87)

Next we write
\[
\frac{m_1' m_2'}{(1 - \sigma m_1 m_2) + (1 - m_1 m_2^2)} = \frac{m_1' m_2'}{(1 - m_1 m_2)(1 - m_1^2)(1 - m_2^2)} \left[ 1 + O(\sqrt{\eta}) \right]
= - \frac{1}{\sqrt{(4 - x^2)(4 - y^2)}} \left[ 1 + O(\sqrt{\eta}) \right]
= \frac{2}{\sqrt{(4 - x^2)(4 - y^2)}} \left[ 1 + O(\sqrt{\eta}) \right],
\]
and so, using similar computations for the second term in (87), we get
\[
\Re \left\{ \frac{m_1' m_2'}{(1 - \sigma m_1 m_2) + (1 - m_1 m_2^2)} \right\} = \frac{2}{\sqrt{(4 - x^2)(4 - y^2)}} \left[ \frac{|x + m_1 + m_2|}{|x + m_1 + m_2|^2} \right] + O(\sqrt{\eta}).
\]
(88)

Combining (87) with (88), we conclude the computation of the variance of $L_N(f, I)$. Using exactly the same computations as in the case $\sigma = 1$, for $\sigma = -1$ we get
\[
\frac{1}{\pi^2} \int_{\mathbb{R}} \int_{\mathbb{R}} \int_{|y| \leq \eta} \int_{|y'| \leq \eta} \Re f(z) \Re g(w) \frac{m_1' m_2'}{(1 - \sigma m_1 m_2) + (1 - m_1 m_2^2)} \Re \left\{ \frac{m_1' m_2'}{(1 - \sigma m_1 m_2) + (1 - m_1 m_2^2)} \right\} dx dy.
\]
We are now left with the case $|\sigma| < 1$. In this case the rhs. of (82) is smooth since $|1 - \sigma m_1 m_2| \geq 1 - |\sigma|$. This implies that, unlike in (88)–(86), we can perform integration by parts without subtracting $f(x)$, $f(y)$. Using (82) once again, for $|\sigma| < 1$ we compute
\[
\frac{1}{\pi^2} \int_{\mathbb{R}} \int_{\mathbb{R}} \int_{|y| \leq \eta} \int_{|y'| \leq \eta} \Re f(z) \Re g(w) \frac{m_1' m_2'}{(1 - \sigma m_1 m_2) + (1 - m_1 m_2^2)} \Re \left\{ \frac{m_1' m_2'}{(1 - \sigma m_1 m_2) + (1 - m_1 m_2^2)} \right\} dx dy + O(\sqrt{\eta}).
\]
(89)
Then we compute
\[
\frac{\sigma m_1^2 m_2^2}{(1 - \sigma m_1 m_2)^2} = \frac{1}{2 \sqrt{(4 - x^2)(4 - y^2)}} \left[ \frac{1 + \sigma^2}{|x + m_1 + \sigma m_2|^2} - \frac{1 - \sigma^2}{|x + m_1 + \sigma m_2|^2} \right],
\]
and so using similar computations for the second term in the rhs. of (189), we conclude that
\[
\Re \left[ \frac{\sigma m_1^2 m_2^2}{(1 - \sigma m_1 m_2)^2} - \frac{\sigma m_1^2 m_2^2}{(1 - \sigma m_1 m_2)^2} \right]
= \frac{(1 + \sigma^2) (2(1 + \sigma^2) - \sigma xy)}{2[(1 - \sigma^2)^2 - (1 + \sigma^2) \sigma xy + \sigma^2 (x^2 + y^2)] (4 - x^2)(4 - y^2)}
- \frac{1}{2}(1 - \sigma^2)^2 \left[ \frac{8 \sigma^2 - 2 \sigma (1 + \sigma^2) \sigma xy + \sigma^2 (x^2 + y^2)^2 (4 - x^2)(4 - y^2)}{2[(1 - \sigma^2)^2 - (1 + \sigma^2) \sigma xy + \sigma^2 (x^2 + y^2)]}ight]
= \frac{1}{2} \partial_x \partial_y \log \left( x - \sigma y \right) \left( \sqrt{4 - x^2} - \sqrt{4 - y^2} \right)
- \frac{1}{2} \partial_x \partial_y \log \left( x - \sigma y \right) \left( \sqrt{4 - x^2} + \sqrt{4 - y^2} \right)
\]
on \eta, y \in [-2, 2] and zero otherwise. This concludes the computation of the variance of \( L_N (f, I) \).

E.2.2. Computation of the variance of \( L_N (f, A_{\text{od}}) \) and \( L_N (f, A_{\text{ul}}) \). In this section we proceed with the computation of the variance of \( L_N (f, A_{\text{od}}) \). The computations for the variance of \( L_N (f, A_{\text{ul}}) \) are completely analogous and so omitted.

Using (77), we start with
\[
\int_{\mathbb{R}} \int_{|z|^\eta} \int_{|y|^\eta} \partial_x f(c(z)) \partial_x g(c(w)) \frac{\sigma m_1^2 m_2^2}{1 - \sigma m_1 m_2}
= -\frac{1}{2} \int_{\mathbb{R}} \int_{|z|^\eta} \int_{|y|^\eta} f(c(x + i\eta)) \partial_x g(c(y + i\eta)) \Re \left[ \frac{\sigma m_1^2 m_2^2}{1 - \sigma m_1 m_2} - \frac{\sigma m_1^2 m_2^2}{1 - \sigma m_1 m_2} \right].
\]

Then we compute
\[
\frac{\sigma m_1^2 m_2^2}{1 - \sigma m_1 m_2} - \frac{\sigma m_1^2 m_2^2}{1 - \sigma m_1 m_2}
= -2i \Re \{ 3 m_2 \}
= -2i \Re \{ 3 m_2 \} - \frac{2 \Re \{ 3 m_2 \}}{x + i\eta + 2 \Re \{ 3 m_2 \} + m_1(1 - \sigma^2 |m_2|^2)}.
\]

Taking the real part we have that
\[
\Re \left[ \frac{\sigma m_1^2 m_2^2}{1 - \sigma m_1 m_2} - \frac{\sigma m_1^2 m_2^2}{1 - \sigma m_1 m_2} \right]
= 23 m_1 \Re \{ 3 m_2 \} (|x|, |y| \leq 2)
- \frac{23 m_2 (\eta_r + 3 m_1 (1 - \sigma^2 |m_2|^2))}{(x + 2 \Re \{ 3 m_2 \} + m_1 (1 - \sigma^2 |m_2|^2))^2 + (\eta_r + 3 m_1 (1 - \sigma^2 |m_2|^2))^2}
= 23 m_1 \Re \{ 3 m_2 \} (|x|, |y| \leq 2) - 1 (\sigma = \pm 1) \sqrt{4 - y^2} \frac{2 \eta y}{(x - \sigma y)^2 + 4 \eta^2}
- \frac{(1 - \sigma^2) \sqrt{(4 - x^2)(4 - y^2)}}{2 \sigma^2 (x^2 + y^2) + (1 - \sigma^2)^2 - x y \sigma (1 + \sigma^2)} \Re \{ 3 m_2 \} (|x|, |y| \leq 2)
\]
in distributional sense.

Similarly, using (77), we readily conclude
\[
\int_{\mathbb{R}} \int_{|z|^\eta} \partial_x f c m^3 = -\int_{-2}^2 f(x) \frac{\sqrt{4 - x^2} (1 - x^2)}{2} \, dx.
\]
Finally, using (171) again, we compute the integral of the last term in the fourth line of (178) and conclude that

\[
\frac{1}{\pi^2} \int_{\mathbb{R}} \int_{|x| \leq n} \partial_x f(z)m^2 \, d\eta \, dx = \int_{-2}^{2} f(x) x \rho_{\text{sc}}(x) \, dx.
\]

This concludes the computation of the variance of $L_N(f, \hat{A}_d)$ and so of Theorems 2.4.

E.3. Independence of linear statistics on different scales: Proof of Theorem 2.13. In the computations above we performed all the analysis for a single scale $N^{-a}$. However, inspecting the proof one can notice that exactly the same computations go through for test functions $f_1$ and $f_2$ living on two different scales $N^{-a_1}, N^{-a_2}$, for $N$-independent $a_1, a_2 \in (0, 1)$ such that $a_1 < a_2$. The case when $a_1 = 0$ is analogous and so omitted. In particular, the test functions $f_i$ are given by

\[
f_i(x) := g_i(N^a(x - E_i)), \quad i = 1, 2.
\]  

The bound of the error term in (178) is completely analogous and it is given by $[N \min_{1, \eta_{a_i}}]^{-1/2}$. The bound of the leading terms follow by estimating the deterministic terms computed in Section E.2 to prove Theorem 2.4. To make this argument clearer we give the explicit bound of the representative term:

\[
\frac{1}{4\pi} \int_{-2}^{2} \int_{-2}^{2} \frac{f_1(x) - f_1(y) (f_2(x) - f_2(y))}{(x - y)^2} \frac{4 - xy}{\sqrt{(4 - x^2)(4 - y^2)}} \, dx \, dy,
\]

which corresponds to $m_1^1 m_2^1 (1 - m_1 m_2)^{-2}$ in (176) after plugging it into the Helffer-Sjöstrand formula as in (178) (see (87)-(88) for the explicit computations that give (99)). For simplicity, we assume $E_1 = E_2 = E$, the general case is analogous. Using the definition of $f_i$ in (194), and the change of variables $z = N^{a_1}(x - E), w = N^{a_2}(y - E)$, we have that

\[
|093| \leq \frac{1}{N^{a_1 + a_2}} \int \int_{\text{supp}(g_1) \cap \text{supp}(g_2)} \left| \frac{(g_1(z) - g_1(w)) (g_2(z) - g_2(w))}{(z N^{-a_1} - w N^{-a_2})^2} \right| \, dw \, dz \lesssim N^{-(a_2 - a_1)},
\]

where we used the smoothness of $g$, that the domain of integrations have size of order one as a consequence of $g_i$ being compactly supported, and that

\[
\frac{4 - xy}{\sqrt{(4 - x^2)(4 - y^2)}} = 1 + O(N^{-a_1})
\]

in the relevant regime.

Appendix F. Proof of (128)

The proof of (128) uses an operator version of Pleijel’s formula that cannot be directly read off from the original paper [Add a reference], so we present a detailed proof here for completeness.

By the spectral theorem for the Wigner matrix $W$ we have that

\[
f(W) = \int_{\mathbb{R}} f(\lambda) \, d\rho(\lambda),
\]

where $d\rho(\lambda)$ is the (projection valued) spectral measure of $W$ and $f \in BV(\mathbb{R})$ is a function of bounded variation. Set

\[
I(z) := \frac{1}{2\pi i} \int_{L(z)} G(w) \, dw,
\]

with $z = x + i\eta$ and $L(z)$ being the contour in Figure 1. Recall that $G(w) = (W - w)^{-1}$ is the resolvent of $W$. We assumed that $W \gtrsim -5$ that holds with very high probability. Then we get

\[
I(z) = \frac{1}{2\pi i} \int_{L(z)} G(w) \, dw = \frac{1}{2\pi i} \int_{L(z)} \frac{1}{\lambda - w} \, dw \, d\rho(\lambda) = \frac{1}{\pi} \int_{\mathbb{R}} \nu(\lambda, z) \, d\rho(\lambda),
\]

with $\nu(\lambda, z) \in (0, \pi)$ being the angle shown in Figure 1. Since for small $\eta$ the angle $\nu(\lambda, z)$ is typically close to $\pi$ or $0$ depending on whether $\lambda < x = \Re z$ or $\lambda > x$, thus $I(z)$ well approximates the spectral projection $\rho(W \leq x)$. In fact, using

\[
\eta \Re G(z) = \int_{\mathbb{R}} \frac{\eta (\lambda - z)}{(\lambda - x)^2 + \eta^2} \, d\rho(\lambda) = \int_{\mathbb{R}} \sin \nu(\lambda, z) \, d\rho(\lambda),
\]

\[
\eta \Im G(z) = \int_{\mathbb{R}} \frac{\eta (\lambda - z)}{(\lambda - x)^2 + \eta^2} \, d\rho(\lambda) = \int_{\mathbb{R}} \cos \nu(\lambda, z) \, d\rho(\lambda),
\]
and
\[ \eta \Im G(z) = \int_{\mathbb{R}} \frac{\eta^2}{(\lambda - x)^2 + \eta^2} \, d\rho(\lambda) = \int_{\mathbb{R}} \sin^2 \nu(\lambda, z) \, d\rho(\lambda), \]
we obtain the identity
\[ \rho([-5, x]) = I(z) + \eta \int_{\mathbb{R}} \gamma G(z) + \int_{\mathbb{R}} g(\lambda, z) \, d\rho(\lambda), \quad z = x + i\eta, \]
with
\[ g(\lambda, z) := \nu(\lambda) - \pi \chi_{(-\pi/2, \pi)}(\nu(\lambda, z)) - \sin \nu(\lambda, z) \cos \nu(\lambda, z), \]
where we also used that \( \chi_{(-5, x)}(\lambda) = \chi_{(-\pi/2, \pi)}(\nu(\lambda, z)). \) A simple calculation shows that the function \( g \) satisfies the bound
\[ |g(\lambda, z)| \leq \sin^2 \nu(\lambda, z). \]
Combining (196) and (201), and using integration by parts, we get
\[ f(W) = \int_{\mathbb{R}} \left[ I(\lambda + i\eta) + \frac{\eta}{\pi} \gamma G(\lambda + i\eta) + \int_{\mathbb{R}} g(s, \lambda + i\eta) \, d\rho(s) \right] \, d\lambda. \]
Then choosing \( \eta = \eta_0, \) with \( \eta_0 \) given below (129), and using (203) for \( f(W) = P(W), \) where \( P(W) \) is defined in (127), we conclude
\[ \langle P(W)A \rangle = \frac{1}{2\pi i} \int_{\Gamma_{K, i0}} \langle G(w)A \rangle \, dw \]
\[ + \sum_{\pm} \eta_0 \frac{1}{\pi} \gamma \left[ \langle \gamma \eta_0 \pm \eta_0 \rangle A \right] + \int_{\mathbb{R}} g(s, \gamma \eta_0 \pm \eta_0 + i\eta_0) \langle \rho(s)A \rangle \right], \]
with \( \gamma \eta_0 \) defined in (124), \( \eta_0 \) in (123), and \( \Gamma_{K, i0} \) in (129), and \( A \) is a deterministic matrix with \( \langle A \rangle = 0 \) and \( \|A\| \lesssim 1. \) For the first term in the r.h.s. of (203) we used the definition of \( I(z) \) in (197).
Finally, using that
\[ \langle \rho G(x + i\eta_0) \rangle \lesssim \frac{\sqrt{\rho(x + i\eta_0)}}{N^{3/2}}, \]
for any \( |x| \leq 5 \) by the local law (40), and that
\[ \left| \int_{\mathbb{R}} g(s, x + i\eta_0) \langle \rho(s)A \rangle \right| = \left| \int_{\mathbb{R}} g(s, x + i\eta_0) \frac{1}{N} \sum_{i} \delta(s - \lambda_i) \langle u_i, Au_i \rangle \, ds \right| \]
\[ \lesssim \frac{1}{N^{1/2}} \sum_{i} \int_{\mathbb{R}} \sin^2 \nu(s, x + i\eta_0) \delta(s - \lambda_i) \, ds \]
\[ \lesssim \frac{\eta_0}{\sqrt{N}} \langle \Im G(x + i\eta_0) \rangle \lesssim \frac{\eta_0 \rho(x + i\eta_0)}{\sqrt{N}}, \]
for any \( |x| \leq 5 \) by (204) we conclude the proof of (128). Here \( \{\lambda_i\}_{i \in [N]} \) are the eigenvalues of \( W \) and \( \{u_i\}_{i \in [N]} \) are the corresponding eigenvectors. Note that to go from the first to the second line we used that \( \langle |u_i, Au_i| \rangle < N^{-1/2} \) by [15, Theorem 2.2] and (202), and to go from the second to the third line we used (200).
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