Variability in high-throughput ion-channel screening data and consequences for cardiac safety assessment

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A B S T R A C T

Introduction: Unwanted drug interactions with ion currents in the heart can lead to an increased pro-arrhythmic risk to patients in the clinic. It is therefore a priority for safety pharmacology teams to detect block of cardiac ion channels, and new technologies have enabled the development of automated and high-throughput screening assays using cell lines. As a result of screening multiple ion-channels there is a need to integrate information, particularly for compounds affecting more than one current, and mathematical electrophysiology in-silico action potential models are beginning to be used for this. Methods: We quantified the variability associated with concentration-effect curves fitted to recordings from high-throughput Molecular Devices IonWorks® Quattro™ screens when detecting block of I Na (HERG), I Ks (KCNQ1/minK) and I Na (CaV1.2), for control compounds used at AstraZeneca and GlaxoSmithKline. We examined how variability propagates through in-silico action potential models for whole cell electrical behaviour, and how confidence intervals on model predictions can be estimated with repeated simulations. Results: There are significant levels of variability associated with high-throughput ion channel electrophysiology screens. This variability is of a similar magnitude for different cardiac ion currents and different compounds. Uncertainty in the Hill coefficients of reported concentration-effect curves is particularly high. Depending on a compound’s ion channel blocking profile, the uncertainty introduced into whole-cell predictions can become significant. Discussion: Our technique allows confidence intervals to be placed on computational model predictions that are based on high-throughput ion channel screens. This allows us to suggest when repeated screens should be performed to reduce uncertainty in a compound’s action to acceptable levels, to allow a meaningful interpretation of the data.

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1. Introduction

Unwanted drug interactions with cardiac ion channels can lead to an increased pro-arrhythmic risk in the clinic (Roden, 2008). It is therefore a priority for safety pharmacology teams to detect these interactions as early as possible during drug development. As such, it is now common for pharmaceutical companies to perform high-throughput screening (HTS) on large numbers of novel compounds, early in development, to detect whether they block cardiac ion channels (Pollard et al., 2010). Discovering that a lead compound carries a high cardiac risk at later stages of development is very costly.

At both AstraZeneca (AZ) and GlaxoSmithKline (GSK), results from HTS assays for multiple cardiac ion channels inform mathematical models for a compound’s action on whole-cell electrophysiology (Davies et al., 2012; Mirams et al., 2011). Simulations are run to predict
whether significant changes to cellular electrophysiology may occur with a given compound, and at which concentrations. Development of compounds that are flagged as having a significant effect can then be de-prioritised, or they can be referred for further safety testing, which is more accurate, but also more expensive.

In this study we characterised the variability associated with the results of HTS assays for cardiac ion channels, and examined how that variability affects predictions that are made for drug-induced changes to whole-cell electrophysiology.

1.1. High throughput screening

The current ‘gold standard’ of electrophysiological screening for ion-channel block involves manual patch clamping of single cells. This method is resource intensive and technically challenging, consequently it has a throughput that is far too low to be of commercial use during early drug discovery, when a large number of compounds need to be screened. To some extent, this problem has been overcome in recent years with the development of dedicated planar array high-throughput electrophysiology platforms. These platforms allow high precision recordings of ion-current inhibition by compounds, while being amenable to high-throughput screening (HTS).

A number of different planar array electrophysiology platforms have been developed in recent years, namely: Qpatch by Sophion Bioscience (Mathes, Friis, Finley, & Liu, 2009), PatchXpress by Molecular Devices (Ly, Shyy, & Misner, 2007), PatchLiner by Nanion Technologies (Polonchuk, 2012), IonFlux by Fluxion (Golden et al., 2011), CytoPatch by Cytocentrics (Stett, Burkhardt, Weber, Van Stiphout, & Knott, 2003) and IonWorks™ HT/Quattro and Barracuda by Molecular Devices. A recent comparison of hERG IC50 values obtained across these platforms with manual patch clamp experiments can be found in Gillie, Novick, Donovan, Payne, and Townsend (2013, Table 2). The Molecular Devices Fluorometric Imaging Plate Reader (FLIPR® Tetra, Schroeder & Neagle, 1996) assay is also used to measure changes in calcium transients (Sullivan, Tucker, & Dale, 1999).

HTS provides a measurement of the concentration-effect curve of ion-channel current inhibition, and has established a good record in detecting ion-channel blockade for a wide variety of compounds and ion channels.

In this study we consider the Molecular Devices IonWorks® Quattro™ assays for detecting block of IKr (hERG), IkNa (NaV1.5), ICat (CaV1.2), IK (KCNQ1/minK) and Ig (Kv4.3/KChIP2.2), and the Molecular Devices FLIPR assay for ICat (CaV1.2). These assays have been used routinely at AZ and GSK for novel compound screening for a number of years.

1.2. Concentration-effect curves

Ion-channel current inhibition is commonly described by concentration-effect curves, such as those seen in Fig. 1. This curve describes how an ‘effect’ or ‘response’ R depends on a ‘dose’ or compound ‘concentration’ [C]. In this case, the peak ionic current following a voltage step is recorded repeatedly, and the proportion of this that remains after addition of a certain concentration (or dose) of a compound is the recorded effect (or response). Such curves are commonly described by a Hill function (Hill, 1910):

$$R(C) = \frac{[C]^{n}}{[C]^{n} + [C_{50}]^{n}}, \text{ or equivalently } R(C) = \left[1 + \left(\frac{C}{C_{50}}\right)^{n}\right]^{-1}. \quad (1)$$

This function of concentration [C], has two parameters: [C_{50}], the half-maximal inhibitory concentration; and the Hill coefficient n.

In Fig. 1 we present concentration-effect curves that have been fitted to the output of a hERG-channel HTS that was performed 120 times (on different screening plates). Only those HTS assays that passed quality-control criteria, i.e. ones in which the positive controls registered an acceptable positive result, are considered in this study. We are therefore considering only the variability in ‘acceptable’ measurements, not across all the measurements that were taken. For discussion of typical quality control criteria used in these assays please see Bridgland-Taylor et al. (2006), Harmer et al. (2008), and Davies et al. (2012). Whilst the concentration-effect curves consistently detect the compound’s blockade of IKr, it is immediately evident that there is substantial variability between individual curves shown in Fig. 1. Our use of the term ‘variability’ is intended to represent the variation in IC_{50} values and Hill coefficients fitted to concentration-effect curves recorded by HTS assays that are carried out on separate occasions. Indeed, HTS is thought to provide more variable results than the ‘gold-standard’ of manual patch clamp, and we seek to characterise this HTS variability in this study.

1.3. Aims

For most novel compounds that are of interest in drug development, repeats of only N = 1 or N = 2 are common. It is therefore difficult to draw any conclusions about the variability associated with screening a particular novel compound. Fortunately, a large number of repeats exist for those compounds used as controls in the assays.

Since AZ and GSK both run positive controls on each IonWorks Quattro or FLIPR Tetra plate, a whole concentration-effect curve is evaluated for a control compound each time any compound is screened. This has led to the accumulation of an unprecedented amount of information on HTS variability, which we have utilised here to determine the statistical distributions of both IC_{50} values and Hill coefficients that describe the concentration-effect curves.

By using in-silico action potential (AP) simulations, AZ and GSK have begun to integrate quantitative information on concentration-effect curves that is gained from a panel of cardiac ion channel screens (Davies et al., 2012; Mirams et al., 2013). The aim of these simulations is to provide a prediction of how a compound is likely to affect the whole cardiac cell, or even the whole heart, early in compound development (Fletcher et al., 2011; Mirams & Noble, 2011). For example, in related work, we show how HTS data can be used to predict the results of a rabbit left-ventricular wedge assay (Beattie et al., 2013—this issue). It has not yet been considered how variability in HTS results, that are taken as inputs into in-silico models, might affect such simulation outputs.

The aims of this article are therefore two-fold: firstly, to quantify the variability in concentration-effect curves produced by HTS, by examining variability in the IC_{50} values and Hill coefficients that
describe them; and secondly, to provide a method for estimating the subsequent variability in the output of in-silico action potential simulations based upon HTS.

2. Methods

2.1. Ion-channel screening

Here we describe the IonWorks and FLIPR platforms in general terms. For details of the experimental protocols that are used at AZ and GSK, please refer to Supplementary material S1.

The IonWorks™ Quattro platform (first described by Schroeder, Neagle, Trezise, and Worley (2003)) has been the mainstay of ion channel electrophysiology HTS at AZ and GSK in recent years. On this platform, the glass electrodes typical of manual patch clamping rigs have been replaced by planar, 96 or 384 well, “chips” (known as the PatchPlate™). Each of the wells contains one hole to which one cell is clamped. Briefly, the PatchPlate™ is placed on the machine at the interface between two separate fluid compartments. The extracellular compartment (above the holes), is loaded with an external solution. The surface below the PatchPlate™ (the intracellular compartment), is perfused with a second solution. A vacuum is used to attach the cells onto the small holes at the base of each well, creating a high resistance seal (∼100 MΩ) between the cell and the edge of the PatchPlate™.

Unlike other methods, this suction is not used to break the cell membrane. Instead, a cell membrane-perforating agent (Amphotericin-B) is introduced into the intracellular compartment allowing access to the intracellular compartment from the underside while the high-resistance seal is maintained. This method, known as ‘perforated patch’, allows many of the intracellular components required for ion channel modulation to be retained (Wood, Williams, & Waldron, 2004). The cell can then be voltage clamped and currents across the membrane are measured by a 48 channel amplifier. Separate measurements can be taken from each well using this method.

An updated method, known as population patch clamping (PPC) was developed by Dale, Townsend, Hollands, and Trezise (2007). This protocol can be utilised on IonWorks™ Quattro and Barracuda machines and allows for more accurate recordings. In PPC mode, each PatchPlate™ is perforated by many holes allowing one amplifier to take recordings from up to 64 different cells. The recorded current is then the mean current from a number of cells. This dramatically improves consistency in the recordings, by helping to overcome some of the inter-cell variability that has been observed with the single-hole approach (Finkel et al., 2006). All of the data discussed in this study were acquired using the IonWorks Quattro in the PPC mode, as described in Supplementary material S1. In brief, cells were placed into wells at a concentration of between 1 and 5 million cells/ml (depending on the HTS target), cells were incubated in the presence of the test compound for 3–5 min before measurements were taken, and compounds were tested up to the limit of solubility.

We also consider the FLIPR screen that GSK has been using to detect blockade of CaV1.2 channels (Sullivan et al., 1999). This screen uses calcium-sensitive dyes to record changes in fluorescence. Intracellular calcium transients occur upon the addition of a depolarising solution, reductions in these transients relative to control wells are used as a proxy for blockade of the L-type calcium current. Details of the protocols that were used are given in Supplementary material S1.3.

2.2. pIC50 and Hill coefficient distributions

Of particular note, in terms of size, is the IonWorks hERG control at AZ, in which the compound Cisapride is used. At the time of writing there had been N = 12,638 replications of this experiment. We therefore use this example to discuss the fitting of the data to particular probability distributions. A histogram of the pIC50 values that have been recorded is shown in Fig. 1. For ease of presentation and intuition we will work with pIC50 values throughout this article. These are calculated from IC50 values using pIC50 = −log10(IC50) with units of Molar for IC50, and therefore log(Molar) for pIC50. Probability plots revealed that the logistic distribution is a good descriptor of pIC50 variability in this dataset, this was true for all of the different HTS control assays (please see Supplementary material S2 for more detail).

To estimate the parameters of the distribution we used maximum likelihood estimates, provided by the MatLab™ statistics toolbox ‘mle’ function (the datasets and fitting code are available from download from http://www.cs.ox.ac.uk/chaste/download). The logistic distribution is fitted to the Cisapride hERG control data, and is shown overlaid in Fig. 1.

Throughout this investigation we therefore assume that pIC50 follows a logistic distribution (pIC50 ∼ Logistic(µ, σ), where µ is the centering parameter and σ is the spread parameter). Note that since the pIC50 is calculated from the IC50 using a logarithm, this implies that IC50 follows a log-logistic distribution (IC50 ∼ LogLogistic(α, β)), where the two distributions are related by α = eθ and β = 1/θ. For the equations of the distributions please see the Supplementary material S2.

We perform a similar characterisation of the variability in Hill coefficients, that were fitted to concentration-effect curves at the same time as pIC50 values. In this case all Hill coefficients appeared to follow a log-logistic distribution, as IC50 values do (Hill = LogLogistic(α, β)). In Fig. 1 we present a histogram of the Hill coefficients that correspond to the same assay as the pIC50 values shown in Fig. 1.

The peaks at 1.0 and 1.1 in Fig. 1 may illustrate unconscious ‘rounding’ by human operators, affecting the values that are logged into a database, whereas no such bias exists for pIC50 values. Also of note is the accumulation of Hill coefficients at 5.0, which was the maximum allowed Hill coefficient in the fitting process.

2.3. Inferring underlying distributions from limited repeats

In order to simplify this section we discuss pIC50 values, exactly the same procedure can be followed for Hill coefficients by substituting α for µ and β for σ, in the following discussion.

When an HTS assay is performed on a novel compound, we will typically have only a limited number of concentration-effect curves, and hence very limited knowledge of the distribution of concentration-effect curve parameters (in contrast to the control cases discussed above, where we have a very good approximation to the true distribution).

We must therefore make some assumptions about a distribution for the novel compounds:

1. When an HTS is repeated a large number of times, µ is equal to the ‘true’ pIC50 value. This is a reasonable approximation since µ is equal to the mean, median and mode of the logistic distribution (and for the Hill coefficient case, with the log-logistic distribution, α is equal to the median). Note this is only really the ‘true’ pIC50 if the HTS assay has no systematic bias resulting in under- or over-reporting of pIC50 values.

2. The deviation parameter (σ) is the same as that of the relevant control assays: i.e. if we repeated the HTS for the novel compound hundreds of times we would fit a distribution with the same deviation (σ) as the relevant control for that HTS assay, but with a different median (µ). As we will see in Section 3.1, this is a reasonable assumption for most of the assays, as different control compounds tend to share similar deviation parameters, and for a first approximation we might expect a novel compound to do the same.

In Supplementary material S3 we provide details of how these assumptions are used in inference calculations to estimate the
underlying distribution of the true \( \text{pIC}_{50} \), from individual runs of the assay. Here we discuss an example of the inference procedure, shown in Fig. 3.

In Fig. 3 we observe how the incorporation of repeat assay results into our inference procedure influences our confidence in the possibilities for the ‘true’ \( \text{pIC}_{50} \). If repeat experiments yield similar values (as will we see, this is suggested by the results in Section 3.1), the effect of including them is to reduce the effective spread of the probability distribution for \( \mu \). We therefore tend to obtain more accurate estimates for the ‘true’ \( \text{pIC}_{50} \) values when multiple experiments are performed, as demonstrated in Fig. 3.

2.4. In-silico action potential modelling

2.4.1. Model

In-silico action potential models are based on the Nobel Prize winning work of Hodgkin and Huxley (1952). These models describe the evolution through time of voltage across a cell membrane that results from differences in intracellular and extracellular ionic concentrations, caused by transmembrane ion currents, which are themselves voltage and time dependent. The first cardiac action potential models were developed over 50 years ago (Noble, 1962), and models now exist for multiple species and types of cardiac myocyte (Noble, Garmy, & Noble, 2012). These cellular models have also been integrated into models of tissue electrophysiology to predict the path of electrical waves up to the whole heart scale (Clayton et al., 2011). We have recently reviewed the use of these biophysical cardiac electrophysiology models in pro-arrhythmic safety testing (Mirams, Davies, Cui, Kohl, & Noble, 2012).

2.4.2. Sampling from distributions

In the results section we consider data from a screen of a number of marketed reference compounds. This screen was performed once, so to emulate the process of repeat HTS experiments we assume the screened values represent the ‘true’ values. If the experiment was to be repeated we assume that the individual sample \( \text{pIC}_{50} \) values and Hills would follow the PDF of those HTS screens, assuming true median values for \( \mu \) to be those given in Table 2 and spread parameters to be as in Table 1. The Bayesian inference scheme is then applied to 1, 2, 4 or 8 data points, drawn at random from these distributions (if one did possess raw data from multiple repeats the Bayesian inference scheme should be applied directly to these values).

Having inferred a probability density function for \( \mu \), as described in Section 2.3, we can then sample an estimate for the ‘true’ \( \text{pIC}_{50} \) which is subsequently used to create a sample concentration-effect curve \( R([C]) \), using Eq. (1).

Then for any drug concentration of interest, the sample concentration-effect curve is used to calculate a new maximal channel conductance using the conductance block model (Eq. (2)). This process is repeated as required for all channels of interest.

We differentiate between the cases where we allow: only \( \text{pIC}_{50} \) to be sampled (and Hill coefficient is assumed to be equal to one) or both \( \text{pIC}_{50} \) and Hill (independently, as discussed in Supplementary material S3); and the inclusion of variability on different numbers of the ion currents (just hERG, or hERG/NaV/CaV or hERG/NaV/CaV/IKs/Ito).

2.4.3. Simulations

A \( \text{pIC}_{50} \) and Hill coefficient are selected, for each channel, using the sampling method described above in Section 2.4.2. These values define a concentration-effect curve for each ion current. We modified the Ten Tusscher and Panfilov (2006) model to allow scaling of the channel conductances in response to drug block, as described in Section 2.4.1 and Eq. (2). We begin with the model in steady state 1 Hz pacing drug-free control conditions. The channel conductance scalings are applied, then regular 1 Hz pacing continues until the model reaches a new steady state (limit cycle). The action potential duration (at 90% repolarization, APD$_{90}$) is then calculated and recorded. Note that any marker of interest could be processed from the simulated action potentials, for example upstroke velocity, or

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**Fig. 2.** Histograms of (a) \( \text{pIC}_{50} \) values and (b) Hill coefficients, from an IonWorks Quattro Cisapride hERG assay. These values were fitted to recordings taken at AZ for \( N = 12,638 \) independent assay runs. Solid red lines denote the distributions that have been fitted to these data as described in the text.
triangulation. Here we have restricted the analysis to APD_{90}, as this is the most widely used marker in computational cardiac safety assessment at present, and has been linked to human clinical Torsadogenic risk (Mirams et al., 2011). The process is repeated for a range of concentrations for each sample of pIC_{50} and Hill coefficient. This ‘workflow’ is depicted in schematic form in Fig. 4.

A new pIC_{50} and Hill coefficient are sampled, and the process shown in Fig. 4 is repeated for the new values. By repeating this process 200 times for each compound, resulting in 200 distinct APD-response curves, we can build up a probability distribution for the outputs. By ignoring the 5 maximum and 5 minimum of the 200 simulated APDs at each concentration, the remaining spread of APDs defines an estimate for the 95% credible interval for the simulated response curve. The simulation takes a few hours on a single processor of a modern desktop PC (the time varies depending on the action-potential model used), but moving the calculations to parallel evaluation on larger computing resources is trivial.

For the interested reader we have made the following resources available: the full control datasets to which distributions were fitted, with the corresponding MatLab script; the Bayesian sampling code; the action potential modelling software; and the scripts for generating the figures presented in this article. These can be downloaded as a ‘bolt-on project’ for the open-source cardiac simulation software Chaste (written to work with v3.1) from http://www.cs.ox.ac.uk/chaste/download.

A brief description of the computational techniques follows. We used a machine readable version of the Ten Tusscher and Panfilov (2006) model, taken from the CellML repository (Lloyd, Lawson, Hunter, & Nielsen, 2008). PyCML was used to convert the CellML format into C++ code, the CellML file being tagged with metadata denoting the conductances of interest (Cooper, Corrias, Gavaghan, & Noble, 2011; Cooper, Mirams, & Niederer, 2011). The equations were solved using the adaptive CVODE solver (Hindmarsh et al., 2011; Cooper, Mirams, & Niederer, 2011). The equations were solved using the adaptive CVODE solver (Hindmarsh et al., 2011; Cooper, Mirams, & Niederer, 2011). The equations were solved using the adaptive CVODE solver (Hindmarsh et al., 2011; Cooper, Mirams, & Niederer, 2011). The equations were solved using the adaptive CVODE solver (Hindmarsh et al., 2011; Cooper, Mirams, & Niederer, 2011). The equations were solved using the adaptive CVODE solver (Hindmarsh et al., 2011; Cooper, Mirams, & Niederer, 2011). The equations were solved using the adaptive CVODE solver (Hindmarsh et al., 2011; Cooper, Mirams, & Niederer, 2011). The equations were solved using the adaptive CVODE solver (Hindmarsh et al., 2011; Cooper, Mirams, & Niederer, 2011). The equations were solved using the adaptive CVODE solver (Hindmarsh et al., 2011; Cooper, Mirams, & Niederer, 2011). The equations were solved using the adaptive CVODE solver (Hindmarsh et al., 2011; Cooper, Mirams, & Niederer, 2011).

### 3. Results

#### 3.1. Screening variability

In this section we present the variability associated with fitting concentration-effect curves to the output of HTS assays. Each of the

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**Table 1**

Distribution of concentration-effect curves for control compounds. This table lists the fitted parameters for the pIC_{50} (logistic) and Hill (log-logistic) distributions for each control compound, for each cardiac ion channel assay.

| Ion channel | Control compound | HTS platform | Performed at | Repeats N | pIC_{50} Logistic | pIC_{50} Log-logistic |
|-------------|------------------|--------------|--------------|-----------|-------------------|----------------------|
| NaV1.5      | Flecainide       | IonWorks     | AZ           | 2307      | 5.235 ± 0.0760    | 0.1035               |
| NaV1.5      | Amitriptyline    | IonWorks     | GSK          | 362       | 5.765 ± 0.1388    | 1.744 ± 0.1983       |
| NaV1.5      | Tetracaine       | IonWorks     | GSK          | 248       | 6.060 ± 0.1459    | 1.530 ± 0.2089       |
| NaV1.5      | GSK-A            | IonWorks     | GSK          | 121       | 5.315 ± 0.2044    | 0.930 ± 0.1529       |
| CaV1.2      | Verapamil        | IonWorks     | AZ           | 369       | 5.571 ± 0.1597    | 0.605 ± 0.1206       |
| CaV1.2      | Nicardipine      | FLIPR        | GSK          | 395       | 7.378 ± 0.2216    | 1.225 ± 0.2386       |
| CaV1.2      | Nifedipine       | FLIPR        | GSK          | 328       | 7.248 ± 0.1856    | 1.179 ± 0.2213       |
| CaV1.2      | Diltiazem        | FLIPR        | GSK          | 244       | 5.240 ± 0.1560    | 0.979 ± 0.2263       |
| hERG        | Cisapride        | IonWorks     | AZ           | 12638     | 6.408 ± 0.1034    | 1.790 ± 0.1784       |
| hERG        | Quinidine        | IonWorks     | GSK          | 120       | 5.625 ± 0.1033    | 1.708 ± 0.1544       |
| hERG        | Pimozide         | IonWorks     | GSK          | 94        | 7.321 ± 0.1914    | 1.586 ± 0.2486       |
| hERG        | Haloperidol      | IonWorks     | GSK          | 84        | 6.852 ± 0.1498    | 1.469 ± 0.2031       |
| hERG        | Verapamil        | IonWorks     | GSK          | 72        | 6.169 ± 0.1464    | 1.429 ± 0.2025       |
| I_{Na}      | XE991            | IonWorks     | AZ           | 325       | 6.217 ± 0.1053    | 1.127 ± 0.1510       |
| I_{Na}      | XE991            | IonWorks     | GSK          | 79        | 5.927 ± 0.1342    | 1.011 ± 0.1837       |
| I_{K}       | GSK-B            | IonWorks     | GSK          | 451       | 7.414 ± 0.1808    | 1.318 ± 0.1677       |
| I_{K}       | Flecainide       | IonWorks     | AZ           | 366       | 4.860 ± 0.0860    | 1.063 ± 0.0862       |
compounds we consider has been screened a large number of times as a positive control at either AZ or GSK.

Each assay generates a histogram of reported pIC50 and Hill coefficients, such as those shown in Fig. 2 for Cisapride in the hERG IonWorks screen. Analogous figures are presented in Supplementary material S4 for each control compound. The probability distributions that were fitted to the datasets for each control compound, as discussed in Section 2.2, are shown in one figure for each assay (Figs. 5 & 6). The individual fit parameters are summarised in Table 1.

Of particular interest in Table 1 is the consistency in the values of the ‘spread’ parameters $\sigma$ and $1/\beta$ for each assay, as we would expect $\mu$ and $\alpha$ to vary between different compounds. Note that the spread of the Hill (log-logistic) distribution shown in Figs. 5 & 6 naturally increases as the median $\alpha$ increases, even with an unchanged scaling $(1/\beta)$ parameter (see Supplement Fig. S4). The values in Table 1 may therefore be a better representation of the differences in the variability associated with the Hill coefficients than Figs. 5 & 6.

Those assays with multiple controls, which are hERG (Fig. 4), CaV1.2 FLIPR (Fig. 4), and I$_{\text{Ks}}$ (Fig. 5), demonstrate similar levels of variability in both pIC50 and Hill coefficients for all of the compounds considered. These results provide evidence that the second assumption we made in Section 2.4.2 about consistency in the spread of pIC50 and Hill coefficients is reasonable for these assays. A possible exception to this is the NaV1.5 IonWorks screen, that shows slightly higher variability in the pIC50 spread associated with different compounds. Ideally data would be acquired on additional NaV1.5 control compounds to characterise this assay further. In the absence of any further information, we keep the same assumption for the CaV1.2 IonWorks and I$_{\text{Ks}}$ screens, for which data were only available for one control compound each.

Figs. 5 & 6 and Table 1 show that the variability associated with Hill coefficients fitted to HTS results is considerable (but the level of variability is consistent between different compounds in a given assay). Such high variability may suggest that there would be less error associated with making the assumption that the Hill coefficient is always equal to one, than there is in taking the Hill coefficient from an HTS assay.

There did not appear to be a correlation between the pIC50 and Hill coefficients recorded in any of the assays (see Supplementary material Fig. S5).

3.2. Simulation variability

In this section we show how the uncertainty in the ‘true’ concentration-effect parameters propagates through an action potential simulation. We performed HTS at AZ for the ion channels of interest for a number of reference compounds, the results are shown in Table 2.

We selected an underlying spread parameter for each assay, shown in Supplementary material Table S1, based on the data shown in Table 1. Where there was a choice to be made between AZ or GSK spread parameters, we chose the AZ ones, as the experiments for this section were performed using those reagents, protocols and machines. We then performed the Bayesian inference method to obtain probability distributions for the ‘true’ pIC50 value and Hill coefficient, as discussed in Section 2.4.2. We sampled from these distributions to explore the range of possible concentration-effect curves that are input into the simulations (as shown in Fig. 4). In practice to apply this method, you should characterise the variability of your own assay as shown above, and use the resulting spread parameters in the inference step.

The results consist of concentration-effect curves for action potential duration (APD$_{90}$), with 95% confidence limits evaluated from the many trajectories of the individual runs, as shown in Tables 3 & 4 and Tables S2–S9. In the following sections we present the variation we expect to be associated with simulation outputs, based on the level of variability in the HTS ion channel screens, for each of the compounds listed in Table 2.

3.2.1. Alfuzosin case study

Table 3 summarises the findings for Alfuzosin when considering the variability associated with pIC50 measurements, but not Hill coefficients. Table 4 summarises the findings when including both sources of variability.

In the Thorough-QT (TQT) study a maximum QTc prolongation of 7.7 ms was observed (Gintant, 2011). In Tables 3 & 4 we see that simulations based on single screens could give misleading results. In some cases small, and in others very large, prolongation is predicted at high concentrations (top right plots in each table), particularly when considering variability associated with Hill coefficient measurements (Table 4).

By repeating the ion channel screens we gain more reliable channel block estimates, and the lower panels exhibit the correct behaviour — mild/moderate prolongation of repolarization.

3.2.2. Dofetilide case study

The remaining simulated concentration-action potential duration (APD$_{90}$) curves for this section are presented in Supplementary material S5. Tables S2 & S3 summarise the findings for Dofetilide.

Dofetilide is a well known QT-prolonging hERG blocker, we see that prolongation occurs at low concentrations of 0.01–0.1 $\mu$M. Since the hERG IC$_{50}$ is in this range, a large amount of variability is introduced at these concentrations. When considering only hERG block (Tables S2 & S3, left columns) the variability is reduced at high concentrations. This is due to the blockade of hERG being close to 100% at 1 $\mu$M and above, regardless of the variation in IC$_{50}$ reported by the HTS assay. In a similar situation, at 100 $\mu$M of Quinidine in Fig. 1, all 120 assay runs reported at least 90% block.

When we consider the results of the other channel assays, variability is introduced at higher concentrations, where those assays’ concentration-effect curves become active (Tables S2 & S3, centre and right columns).
In particular, the next IC_{50} to be encountered is that of CaV1.2 (Table 2); and as we might expect blockade of the L-type calcium current at concentrations close to 100 μM begins to cause significant shortening in some of the runs, and correspondingly large confidence intervals on the simulation output.

3.2.3. Lacosamide case study

Lacosamide is an example of a QT-shortening compound. It has a very low hERG affinity, and a significant CaV1.2 affinity. In the TQT study Lacosamide showed a shortening of around 6.3 ms (Gintant, 2011).

Tables S4 & S5 summarise the simulation results for Lacosamide. As we might expect, including assay variability for hERG has almost no effect, as the concentrations we are considering (up to 100 μM) are never close to any of the possible ‘true’ hERG IC_{50} values.

The majority of variability is associated with the CaV1.2 IC_{50} value, and accordingly the additional variability introduced when considering I_{h0} and I_{to} is negligible. This example shows how it is possible, if unlikely, to get very strange behaviour due to HTS variability, in 0.5–1% of cases with just one assay repeat (N = 1), the simulations predict prolongation instead of shortening. In this case we would choose to screen CaV1.2 multiple times, as this has the largest benefit in terms of constraining the confidence intervals.

3.2.4. Nilotinib case study

Tables S6 & S7 summarise the findings for Nilotinib. In our screens Nilotinib is not a strong blocker of any current, and simulations suggest a mild prolongation, or even a possible shortening at high concentrations and low numbers of repeats. By increasing the number of assay repeats, we would consistently predict prolongation.

In the TQT study Nilotinib was a prolonger with 15.8 ms of QTc prolongation being observed (Gintant, 2011). By examining this compound’s FDA pharmacology review, we find that hERG IC_{50} = 0.13 μM was reported, substantially lower than our screening result (Table 2).

3.2.5. Tolterodine case study

The assays suggested that Tolterodine is a strong hERG blocker (Table 2) that we might expect to have a similar profile to Dofetilide. Tables S8 & S9 summarise the simulation findings for Tolterodine. The

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**Fig. 5.** Probability density functions for: (left) pIC_{50} values; (right) Hill coefficients, as fitted to the distributions resulting from large numbers of repeat screens with control compounds.
difference in the profiles shown (ever-longer prolongation at higher concentrations for Tolterodine but not Dofetilide) is explained by the fact that there is no CaV1.2 block to stabilise the hERG prolongation at higher concentrations, and instead blockade of I\(_{Ks}\) and I\(_{to}\) can be introduced.

This potentially risky profile would suggest repeating most of the screening panel multiple times in order to minimise uncertainty in prolongation predictions.

Table 2
Results of a screen for a number of compounds with varying ion-channel blocking profiles. These screens were performed once for each ion channel at AZ. Where 50% block was not achieved, we have performed a least-squares fit of a concentration-effect curve to the original data, with a minimum allowed pIC\(_{50}\) of 0, a minimum allowed Hill coefficient of 0.5, and a maximum of 5.00.

| Compound | hERG | Na\(_{V1.5}\) | Ca\(_{V1.2}\) | KCNQ1/ minK | Kv4.3/ KChIP2.2 |
|----------|------|-------------|-------------|-------------|-----------------|
|          | pIC\(_{50}\) | Hill | pIC\(_{50}\) | Hill | pIC\(_{50}\) | Hill | pIC\(_{50}\) | Hill | pIC\(_{50}\) | Hill |
| Alfuzosin | 4.66 | 0.92 | 3.69 | 0.71 | 3.75 | 1.64 | 3.95 | 0.50 | 3.04 | 0.50 |
| Dofetilide | 6.85 | 1.88 | 3.46 | 0.79 | 3.77 | 5.00 | 3.65 | 0.50 | 0.00 | 1.00 |
| Lacosamide | 0.00 | 1.00 | 3.34 | 0.72 | 4.33 | 0.44 | 3.62 | 0.50 | 0.00 | 1.00 |
| Nilotinib | 4.20 | 0.53 | 2.97 | 0.71 | 3.67 | 0.67 | 3.39 | 0.50 | 2.48 | 0.72 |
| Tolterodine | 6.89 | 1.58 | 5.24 | 1.08 | 0.00 | 1.00 | 4.10 | 0.60 | 4.93 | 1.07 |

Fig. 6. Probability density functions for: (left) pIC\(_{50}\) values; (right) Hill coefficients, as fitted to the distributions resulting from large numbers of repeat screens with control compounds.

4. Discussion

The variability that we have observed in the high-throughput screens (HTS) could be caused by a large number of factors, including (but not limited to):

- Noise in recording or ‘instrument error’.
- Variability in the number and condition of cells in each well.
- Variability in the concentration and differences in the condition/batch of the compound added.
- Variability in the temperature at which the recordings were taken; temperature is not fixed at a physiological 37° in the screens we have considered.

Firstly, we characterised variability in pIC\(_{50}\) values and Hill coefficients resulting from HTS assays by examining controls that have been repeated a large number of times. pIC\(_{50}\) values fitted to data from an IonWorks or FLIPR screen are described by a logistic distribution, and Hill coefficients follow a log-logistic distribution. We observed moderate variability in pIC\(_{50}\) values (Table 1) that is consistent with a \(\pm \frac{1}{2}\) log\(_{10}\) unit ‘rule of thumb’. There is a larger amount of variability associated with Hill coefficients.

There can be such a high level of variability associated with Hill coefficients that it is not clear that including the Hill coefficient from
a HTS adds useful information: there may be less error introduced by assuming that the Hill coefficient associated with each concentration-effect curve \( n \) in Eq. (1) is equal to one. This is something we are investigating in parallel work (Beattie et al., 2013–this issue).

The use of in-silico action potential models to integrate information from HTS rests on the fact that the average result of a HTS is giving the ‘correct answer’. Careful analysis should be made to ensure that these screens are not consistently over- or under-reporting ion channel blockade, by comparing results with manual patch clamping.

We would like other companies and device manufacturers, that have repeated screens large numbers of times, particularly with other machines and for other ion channels, also to publish their datasets. The variability that is associated with different assays, reagents and protocols could then be characterised, and the most reliable screening methods selected for use.

We also characterised the resulting uncertainty in whole-cell responses via simulating the action potential duration based on the uncertainty in the screening data (that the simulation uses as inputs). This was done by performing a Bayesian inference step: inferring the distribution of likely ‘true’ concentration-effect parameters, from a number of sample assay runs, by assuming that the variability level for a novel compound is similar to that of the relevant control assay compounds. We then sampled possible concentration-effect curves for each channel and performed repeated simulations to build up a probability distribution for the resulting action-potential-duration vs. concentration curves.

The approach we have proposed does not rule out the consideration of other sources of variability, for example due to stochasticity of ion channel gating (Dangerfield, Kay, & Burrage, 2012), or inter-individual variability (Davies et al., 2012), and could be considered in addition to these.

The choice of number of screening repeats to perform initially depends upon the expense of the screen, the level of accuracy that is required, and the time and cost involved in repeating screens to reduce uncertainty at a later date. An organisation could decide to perform each screen once, analyse uncertainty using the proposed method, and perform additional screens as required; whilst another organisation could choose to perform each screen four times, or more, to minimise uncertainty from the outset.

Whatever strategy is adopted, our methodology provides a means to determine whether a particular screen is introducing large amounts of

| N  | hERG variability | hERG/NaV1.5/CaV1.2 variability | hERG/NaV1.5/CaV1.2/\(I_{Ks}/I_{to}\) variability |
|----|------------------|--------------------------------|-----------------------------------------|
| APD\(_{90}\) (ms) | ![Graph](image1) | ![Graph](image2) | ![Graph](image3) |
| 1   | ![Graph](image4) | ![Graph](image5) | ![Graph](image6) |
| 2   | ![Graph](image7) | ![Graph](image8) | ![Graph](image9) |
| 4   | ![Graph](image10) | ![Graph](image11) | ![Graph](image12) |
| 8   | ![Graph](image13) | ![Graph](image14) | ![Graph](image15) |

Table 3
Uncertainty in concentration-effect curves for action potential duration under the action of Alfuzosin, when considering ion-channel assay variability in pIC\(_{50}\) values (not in Hill coefficients), for various numbers of repeats. Each plot displays action potential duration, APD\(_{90}\), as a function of concentration. Black lines represent simulation results for each set of sampled concentration-effect inputs, the red line denotes the result when using the numbers reported by the assay directly, with 95% credibility intervals imposed as described in the text.

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uncertainty, and whether the uncertainty in whole-cell behaviour predictions is at an acceptable level. This is a basic requirement for sensible interpretation of simulation results, and therefore a pre-requisite for the possible future use of simulation in reduction and replacement of animal-based safety tests.

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Appendix A. Supplementary data

Supplementary data for this article can be found online at http://dx.doi.org/10.1016/j.vascn.2013.04.007.

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