OGLE-2012-BLG-0950Lb: THE FIRST PLANET MASS MEASUREMENT FROM ONLY MICROLENS PARALLAX AND LENS FLUX

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ABSTRACT

We report the discovery of a microlensing planet OGLE-2012-BLG-0950Lb with a planet/host mass ratio of \( q \approx 2 \times 10^{-4} \). A long term distortion detected in both MOA and OGLE light curve can be explained by the microlens parallax due to the Earth’s orbital motion around the Sun. Although the finite source effect is not detected, we obtain the lens flux by high resolution Keck AO observation. Combining the microlens parallax and the lens flux reveal the nature of the lens: a planet with mass of \( M_p = 35^{+17}_{-9} \, M_\odot \) is orbiting around an M-dwarf with mass of \( M_{\text{host}} = 0.56^{+0.12}_{-0.16} \, M_\odot \) with a planet-host projected separation of \( r_1 = 2.7^{+0.6}_{-0.7} \, \text{kpc} \) located at \( D_L = 3.0^{+0.9}_{-1.1} \, \text{kpc} \) from us. This is the first mass measurement from only microlens parallax and the lens flux without the finite source effect. In the coming space observation-era with Spitzer, K2, Euclid, and WFIRST, we expect many such events for which we will not be able to measure any finite source effect. This work demonstrates an ability of mass measurements in such events.

Key words: gravitational lensing: micro – planetary systems

1. INTRODUCTION

Gravitational microlensing is a technique by which planets can be detected without measurements of light from the host star (Mao & Paczynski 1991; Gould & Loeb 1992; Gaudi 2012). Microlensing can detect planets that are difficult to detect by other methods such as planetary systems in the Galactic Bulge (e.g., Batista et al. 2014), planets around late M-dwarfs or brown dwarfs (Bennett et al. 2008; Sumi et al. 2016), and even free floating planets not hosted by any stars (Sumi et al. 2011). Compared to other techniques, microlensing is sensitive to Earth mass planets (Bennett & Rhie 1996) orbiting just outside of the snow line where the core accretion theory (Ida & Lin 2004) predicts is the most active planet formation region. Microlensing observations so far have revealed a population of planets beyond the snow line (Gould et al. 2010; Sumi et al. 2010; Cassan et al. 2012; Shvartzvald et al. 2010; Suzuki et al. 2016). Suzuki et al. (2016) finds a steeper slope with \( dN/d \log q \approx q^{-0.9} \) and a break (and possible peak) in the mass ratio function at \( q \approx 1.0 \times 10^{-4} \). We are capable of studying the distance distribution of planets in our Galaxy via microlensing. Penny et al. (2016) suggests the possibility of a lack of planets in the Galactic bulge. The detection of extra solar planets by gravitational microlensing presents a number of challenges.
First, gravitational microlensing is an extremely rare phenomenon with a probability of one per one million stars, and a planetary deviation lasts for only hours or a few days. For these reasons, microlensing observations for exoplanets are conducted toward the Galactic Bulge, the most crowded field in our Galaxy. Whereas hundreds of planets are detected by the radial velocity (RV) method (Butler et al. 2006; Bonfils et al. 2013) and thousands of planetary candidates are detected by the Keplter telescope (Borucki et al. 2010) to date, the microlensing method has been used to detect about 50 exoplanets so far.

Several survey groups have been conducting high cadence survey observations using their telescopes with wide FOV cameras in different time zones. The Microlensing Observations in Astrophysics (MOA; Bond et al. 2001; Sumi et al. 2003) group uses the 2.2-deg² FOV MOA-cam3 (Sako et al. 2008) CCD camera mounted on the 1.8 m MOA-II telescope at the Mt. John University Observatory in New Zealand and alerts the community to about 600 microlensing events per year. The Optical Gravitational Lensing Experiment group (OGLE; Udalski 2003) upgraded their camera to the 1.4-deg² FOV OGLE-IV camera in 2010 (Udalski et al. 2015a) and discovered more than 2000 microlensing events per year in the last few years with the camera mounted on the 1.3 m Warsaw telescope at the Las Campanas Observatory, Chile. The Wise observatory group in Israel also conducts microlensing observations (Gorbikov et al. 2010). In 2015, the Korean Microlensing Telescope Network (KMTNet; Kim et al. 2016) also started their survey observations. Now, the equipment requirements for second-generation microlensing surveys (Gaudi et al. 2009; Gaudi 2012) are fulfilled and the number of planet detections will be increasing over the next few years.

Measuring the mass of a lens $M_L$ and the distance to the lens system $D_L$ is challenging. There are three observables in microlensing which can yield a mass–distance relation of the lens system: the angular Einstein radius $\theta_E$, the microlens parallax $\varpi_L$, and the lens flux. The first two of these yield each mass–distance relation by combining the following relationship between them:

$$M_L = \frac{\theta_E}{\kappa \varpi_L}$$  \hspace{1cm} (1)

with the definitions of $\varpi_L$, $\varpi_L \equiv \varpi_{rel}/\theta_E$, where $\kappa$ is a constant and $\varpi_{rel} \equiv au(1/D_L - 1/D_S)$. One can calculate the mass and distance of the lens system if we can measure any two of these quantities. We can measure $\theta_E$ in the following manner. Included in most models explaining planetary microlensing light curve data is the source star radius in units of $\theta_E$: $\rho \equiv \theta_s/\theta_E$. By estimating the angular radius of the source star, $\theta_s$, by an analysis of the source star’s color and magnitude, and using our modeled value of $\rho$, we arrive at an estimate of $\theta_E$. Microlens parallax can be observed only in relatively rare events and lens flux measurements need follow-up observations with high resolution imaging by an 8-m class telescope. Therefore, only half of planetary events published so far are detected with lens mass measurements and masses of the other half planetary systems are just given their probability distributions by a Bayesian analysis (e.g., Beaulieu et al. 2006; Bennett et al. 2014; Koshimoto et al. 2014; Skowron et al. 2015).

In the microlensing planetary events published so far, there are events with the mass measurements from the angular Einstein radius and microlens parallax (e.g., Bennett et al. 2008; Gaudi et al. 2008; Muraki et al. 2011), from the angular Einstein radius and the lens flux (e.g., Bennett et al. 2006; Batista et al. 2015; Bennett et al. 2015), and from all three relations (e.g., Dong et al. 2009; Bennett et al. 2010; Beaulieu et al. 2016; Bennett et al. 2016), but events with mass measurement from only microlens parallax and the lens flux have not been published to date. This is simply because the angular Einstein radius is observed much more commonly than microlens parallax, as mentioned above. However, it has been possible to measure precise microlens parallax by observing simultaneously from space and ground, thanks to the Spitzer microlensing campaign (Calchi Novati et al. 2015; Udalski et al. 2015b; Yee et al. 2015; Zhu et al. 2015). Also, K2 campaign 9 (K2C9), started in 2016 April, has surveyed the Galactic Bulge for three months to date. By combining K2C9 data and ground-based survey data, it is expected to measure microlens parallax for more than 120 events (Henderson et al. 2015). These next-generation space- and ground-based simultaneous observations for microlensing can measure microlens parallax almost regardless of the event timescale. Microlens parallax should become a more common observable, rather than the angular Einstein radius in the coming next generation, thus the mass measurement without the angular Einstein radius should be important (Yee 2015).

This paper reports an analysis of a microlensing planetary event OGLE-2012-BLG-0950, which is the first event where a mass measurement is possible from only the measurements of the microlens parallax and lens flux. The survey observations of this event are described in Section 2. Section 3 explains our data reduction procedure. Section 4 shows our modeling results. We show the constraint on the angular Einstein radius by the source angular radius derived from the color and light curve modeling in Section 5. In Section 6 we describe our Keck observations, the constraints on the excess flux, and calculate the probability of the contamination to the excess flux. In Section 7, we derive the lens properties by combining microlens parallax and the lens flux. Finally, Section 8 discusses and concludes the results of this work.

2. OBSERVATIONS

Microlensing event OGLE-2012-BLG-0950 was discovered and alerted by the OGLE Early Warning System (EWS) on 2012 June 21 (HJD’ ≡ HJD—2450000 ~ 6100) as a new event located at (R.A., Decl.)$_{2000}$ = (18:08:04.62, −29:43:53.7) or (l, b) = (1°765, −4°634). Another survey group, MOA, independently found the event and alerted that as MOA-2012-BLG-527 on 2012 August 9. The observations by OGLE were conducted on the I-band and V-band and the observations by MOA were conducted by the custom MOA-Red filter, which is similar to the sum of the standard Cousins R- and I-band filters. MOA also observed the event in the I-band using the B&C telescope, a 61 cm telescope for follow-up observation at the same site. The observed light curve is shown in Figure 1.

The anomaly part of this event appeared as a small dip around HJD’ ~ 6149, mainly in the MOA data. The MOA-II telescope observed the anomaly with the 47 minute cadence as the regular survey mode. Because the anomaly was very short, ~1 day, and started after the last OGLE observation, we could
not increase the cadence nor issue the anomaly alert in a timely manner. Nevertheless, the normal cadence is enough to reveal the perturbation caused by the planet. The OGLE data with a cadence of once per night until the anomaly, are also very important for the characterization of this event. In particular, the OGLE data shows us that the dip had not started by $\text{HJD} = 6147.6$, had commenced by $\text{HJD} = 6148.6$, and had almost ended by $\text{HJD} = 6149.6$.

This event does not cross any caustic curves and, unfortunately, MOA could not obtain data on $\text{HJD} = 6148$, which corresponds to the start of the anomaly, owing to bad weather. Because of these factors, we have no data on a steep gradient of magnification in this event, thus we cannot detect a significant finite source effect. In addition, we took AO images of the target in the year following the discovery, using the Keck telescope. We describe the details of the Keck observations and the analysis in Section 6.

3. DATA REDUCTIONS

Our data-sets for the modeling below consist of 1275 OGLE $I$-band data points, 81 OGLE $V$-band data points, 6324 MOA-Red data points, and 382 B&C $I$-band data points.

The OGLE data were reduced by the OGLE Difference Image Analysis (DIA) photometry pipeline (Udalski 2003). The centroid of the cataloged star near the event, which is used for PSF photometry in the standard OGLE pipeline, is significantly different from that of the actual event on the difference image. So, we rerun the PSF photometry with the real centroid for the event manually, to obtain more accurate photometry.

The images taken by the MOA-II telescope and the B&C telescope were reduced by the MOA DIA pipeline (Bond et al. 2001). In the crowded stellar field images of the Galactic bulge, the precision of photometry is very sensitive to seeing. We found a systematic photometry bias correlated with the seeing value in the MOA-Red data. We reduced this systematic error by modeling it with a polynomial of seeing and airmass by using the baseline:

$$F_{\text{corr}} = a_0 + a_1 \text{JD} + a_2 \text{airmass} + a_3 \text{airmass}^2 + a_4 \text{seeing} + a_5 \text{seeing}^2 + a_6 \tan z \cos \phi + a_7 \tan z \sin \phi + a_8 \text{airmass} \tan z \cos \phi + a_9 \text{airmass} \tan z \sin \phi + a_{10} \text{airmass} \tan z \cos \phi \text{seeing} + a_{11} \text{airmass} \tan z \sin \phi \text{seeing}^2 + a_{12} \text{airmass} \tan z \cos \phi \text{seeing}^2$$

where $z$ and $\phi$ are the elevation angle and parallactic angle of the target included to correct the differential refraction,
respectively. $F_{\text{cor}}$ is the additional flux for the correction and the corrected flux $F_{\text{new}}$ is $F_{\text{new}} = F_{\text{cor}} + F_{\text{old}}$, where $F_{\text{old}}$ is the original flux from the DIA pipeline. In the resulting photometry, the $\chi^2$ goodness-of-fit value for the time series of baseline is improved by $\Delta \chi^2 \approx 0.07$ per data point.

The relative error of data points given by the photometry code are robust for a given instrument. However, it is known that the absolute values of uncertainty are underestimated in such stellar crowded fields for various reasons in general. Thus, we empirically normalize the errors in each data-set to estimate the proper uncertainties of fitted model parameters. We used the formula presented in Yee et al. (2012) for normalization, $\sigma_i' = \sqrt{\sigma_i^2 + \epsilon_{\text{min}}^2}$, where $\sigma_i$ is the original error of the $i$th data point in magnitudes, and the parameters for normalization are $k$ and $\epsilon_{\text{min}}$. $k$ and $\epsilon_{\text{min}}$ are adjusted so that the cumulative $\chi^2$ distribution as a function of the number of data points sorted by each magnification of the preliminary best-fit model is a straight line of slope 1. By including $\epsilon_{\text{min}}$, we can correct the error bars at high magnification, which can be affected by flattening errors. But we found unusually large $\epsilon_{\text{min}}$ values for the OGLE I and MOA-Red data (0.02 and 0.09, respectively) and the deviations from a straight line in cumulative $\chi^2$ distribution mainly arose from baseline data points, i.e., not from high magnification data points as expected. Thus, it is not reasonable to normalize errors with these values and we adopt $\epsilon_{\text{min}} = 0$ for these two data-sets. This may indicate that there are some low-level systematics in the light curve. We apply $\epsilon_{\text{min}} = 0$, $k = 1.364$ to OGLE I, $\epsilon_{\text{min}} = 0$, $k = 1.576$ to OGLE V, $\epsilon_{\text{min}} = 0$, $k = 0.907$ to MOA-Red, and $\epsilon_{\text{min}} = 0.0061$, $k = 1.021$ to B&C I. The normalization factors applied for the OGLE I and V data are consistent with those given by Skowron et al. (2016a). Note, that the final best fit model parameters are consistent with the preliminary model parameters before the error normalization. Thus, this procedure of the error normalization does not affect our main result. The parameters of these data sets are also shown in Table 1.

### 4. MODELING

Here, we present and compare the results of our light curve modeling, assuming a standard binary lens and also adding the effects of parallax. We fit the light curves using a Markov Chain Monte Carlo (MCMC) approach (Verde et al. 2003), with magnification calculations from the image centered ray-shooting method (Bennett & Rhie 1996; Bennett 2010).

#### 4.1. Standard Binary Lens

In the case of a point lens, the magnification map on the source plane is circular symmetric around the lens. In the point source point lens case, we can characterize the microlensing light curve with the time of the closest approach of the source to the center of mass of the lens, $t_0$, the minimum impact parameter $u_0$ at $t_0$, and the Einstein radius crossing time (or timescale) $t_{\text{E}} = \theta_E/\mu_{\text{rel}}$, where $u_0$ is in units of $\theta_E$ and $\mu_{\text{rel}}$ is the lens-source relative proper motion. When the lens has a companion, its gravity distorts the magnification map and creates the closed curves called caustics, where the magnification is infinite. In this case, three parameters are added to the fitting parameters above; the mass ratio of two lenses $q$ and their angular separation normalized by $\theta_E$, $s$, which determine the shape and location of the caustics, and the source trajectory with respect to the binary lens axis, $\alpha$, which determines the direction of a one-dimensional slice of the distorted magnification map. When a source star crosses a region with a steep gradient near the caustics in the magnification map, we can observe the finite source effect. Because source stars of most binary lens events cross such regions, we include the source size $\rho \equiv \theta_s/\theta_E$ as a fitting parameter for a binary lens model. With the magnification variation against time, $A(t, x)$, which is defined in terms of the above parameters $x = (t_0, u_0, t_E, q, s, \alpha, \rho)$, we can linearly fit

$$F(t) = f_x A(t, x) + f_b$$

(3)
to a data set and obtain the instrumental source flux $f_x$ and the instrumental blending flux $f_b$ for every telescope and pass-band.

We adopt a linear limb-darkening law with one parameter, $u_t$. According to González Hernández & Bonifacio (2009), we estimate the effective temperature, $T_{\text{eff}} \sim 5500$ K from the source color which is discussed in Section 5, and assume the solar metallicity. With $T_{\text{eff}} = 5500$ K and assuming surface gravity $\log g = 4.0$ cm s$^{-2}$ and microturbulence parameter $\xi = 1.0$ km s$^{-1}$, the limb-darkening coefficients selected from Claret (2000) are $u_{V} = 0.5470$ for OGLE I and B&C I, $u_{V} = 0.7086$ for OGLE V, and $u_{\text{MOA-Red}} = 0.5895$ for MOA-Red, which is the average of the standard I and R filters. Therefore, we used the $u_{I}$ for OGLE I and B&C I, the $u_{V}$ for OGLE V and the mean of the $u_{I}$ and $u_{R}$, $0.5895$ for MOA-Red, the filter which has the range of both the standard I and R filters. Although the best estimated value of $T_{\text{eff}}$ and the limb-darkening coefficients depend on the source magnitude in each model, we keep using the fixed values. However, this does not affect our final result because the finite source effect is very weak in the light curve, as mentioned below. The limb-darkening coefficients we adopt are also shown in Table 1.

We show the parameters of the best-fit models of our standard binary lens modeling in Table 2, where the uncertainties shown are from the 16th/84th percentile values of the stationary distributions given by MCMC. We find a degeneracy between the close model of $s < 1$ and the wide model of $s > 1$ with $\Delta \chi^2 \approx 0.7$. There is a well-known degeneracy in high magnification microlensing events between lens systems with similar mass ratios, but separations $s$ and $1/s$. In microlensing events suffering this degeneracy, the source star passes close to the central caustic which has, in each of the degenerate solutions, a similar shape. However, in this event, the close/wide degeneracy has a different nature, in terms of the caustic geometry. A single resonant caustic is seen in the wide model with $s = 1.007$, while the caustic curves are separated into central caustic and planetary caustics in the close model with $s = 0.890$. As seen in Figure 2, it is understood that the gradients of magnification on the source trajectories are similar in both models, although the caustic shapes are different. The mass ratios are $q \approx 2 \times 10^{-4}$ in both models, indicating the companion has a planetary mass. We find that the

| Data set  | Number of data | $k$  | $\epsilon_{\text{min}}$ | $u_0$  |
|----------|----------------|------|-------------------------|--------|
| OGLE I   | 1275           | 1.364| 0                       | 0.5470 |
| OGLE V   | 81             | 1.576| 0                       | 0.7086 |
| MOA-Red  | 6324           | 0.981| 0                       | 0.5895 |
| B&C I    | 382            | 1.017| 0.00611                 | 0.5470 |
finite source effect is weak and the $\rho$ value is consistent with $\rho = 0$ at the 1σ level. Because a larger $\rho$ value reduces the dip depth in the light curve and does not explain the data, we can place an upper limit on $\rho$.

### 4.2. Parallax

In long timescale microlensing events, such as this one, the effect of Earth’s orbital motion around the Sun may be detectable (Gould 1992; Alcock et al. 1995). This effect is expressed by the microlens parallax vector $\pi_E = (\pi_{E,N}, \pi_{E,E}) = \pi_E \mu_{rel}/\mu_{rel}$ (Gould 2000a). Here, $\pi_{E,N}$ and $\pi_{E,E}$ are the north and east components of $\pi_E$, respectively, whose direction is the same as lens-source relative proper motion. The magnitude $\pi_E \equiv 1 au/\mu_{E}$ is defined by 1 au relative to the Einstein radius projected onto the observer plane $\pi_E = R_{E} D_{L}/D_{S} - D_{L}$.

We show our best-fit parallax models in Figures 1, 2, and Table 2. We found each close and wide solutions has an additional degeneracy between $u_0 > 0(+)$ and $u_0 < 0(-)$. These four degenerate solutions have $\Delta \chi^2 < 4$. The parameters of the degenerate models are consistent with each other to within 1σ error, except for $s$, $\alpha$, and $u_0$. The $\chi^2$ difference between the standard models and the parallax models is significantly large, $\Delta \chi^2 > 110$ for 2 dof difference.

It is known that low-level systematics in the baseline sometimes mimic a high order signal. We therefore check whether the $\Delta \chi^2$ contributions come from where we theoretically expect. The top inset in Figure 3 shows the cumulative distribution for $\Delta \chi^2$ between the Standard close model and the parallax close + model as a function of time. Positive $\Delta \chi^2$ values indicate that the parallax model is favored over the standard model. We find $\Delta \chi^2 \sim 90$ comes from the data during the main peak of the event in 2012 in both MOA and OGLE as expected, and $\Delta \chi^2 \sim 25$ comes from the data in 2013, the next year. The right panel of Figure 1, which is a zoom of the 2013 data, shows slight differences among the models, i.e., the parallax models have the magnification of 1.05 in the start of 2013 while the standard models have 1.0. The bottom panel in Figure 3 shows binned residuals in bins 25 days wide. This shows the clear long term deviation from the standard model. The binned data in the first half of 2013 are mostly above the standard model in both MOA and OGLE, whereas those in the other years are not, which can be well explained by the parallax model. Because these long term distortions are consistent in both MOA and OGLE, we conclude that these long term signatures are real and they are better explained by the parallax models compared to the standard models.

Note. A superscript/subscript indicates the difference parameter’s 84/16 percentile from the best-fit value, respectively.

### Table 2

| Model | $t_0$ (HJD*) | $u_0$ (days) | $q$ (10$^{-4}$) | $s$ (rad) | $\alpha$ (rad) | $\rho$ (10$^{-3}$) | $\pi_{E,N}/\pi_{E,E}$ | $\pi_{E,E}/\pi_{E,E}$ | $\pi_{E}/\pi_{E}$ | $P$ (10$^2$) days | $I_0$ (mag) | $\chi^2$ | dof |
|-------|--------------|--------------|-----------------|----------|---------------|-------------------|-----------------------|----------------------|----------------|-----------------|-------------|--------|-----|
| Standard close | 6151.58 | 66.4 | 0.104 | 2.3 | 0.890 | 5.098 | 4.3 | ... | ... | ... | ... | 19.26 | 8168.8 | 8047 |
| $\sigma$ | 0.03 | 1.5 | 0.003 | +0.6 | -0.3 | 0.008 | -0.009 | <7.0 | ... | ... | ... | 0.03 | ... | ... |
| Parallax close | 6151.47 | 68.4 | 0.101 | 1.9 | 0.895 | 5.081 | 0.0 | 0.14 | -0.19 | 0.24 | ... | 19.29 | 8053.2 | 8045 |
| $\sigma$ | 0.03 | 1.5 | 0.003 | +0.1 | -0.007 | 0.006 | <3.2 | +0.11 | 0.02 | -0.04 | ... | 0.03 | ... | ... |
| close | 6151.48 | 68.5 | -0.100 | 1.9 | 0.893 | 1.195 | 0.0 | 0.08 | -0.17 | 0.19 | ... | 19.30 | 8055.8 | 8045 |
| $\sigma$ | 0.03 | 1.6 | 0.003 | +0.5 | +0.006 | -0.007 | <3.0 | +0.12 | 0.02 | -0.02 | ... | 0.03 | ... | ... |
| close + | 6151.49 | 68.7 | 0.101 | 1.9 | 1.004 | 5.083 | 0.0 | 0.17 | -0.20 | 0.26 | ... | 19.29 | 8054.7 | 8045 |
| $\sigma$ | 0.03 | 1.4 | 0.003 | +0.6 | +0.008 | 0.006 | <3.0 | 0.08 | 0.02 | 0.06 | ... | 0.03 | ... | ... |
| wide | 6151.48 | 67.8 | -0.101 | 1.9 | 1.004 | 1.195 | 0.0 | 0.09 | -0.17 | 0.19 | ... | 19.29 | 8057.3 | 8045 |
| $\sigma$ | 0.02 | 1.4 | 0.003 | +0.6 | +0.009 | -0.003 | <3.1 | +0.10 | 0.02 | 0.07 | ... | 0.03 | ... | ... |
| Xallarap close + | 6151.50 | 67.2 | 0.103 | 2.0 | 0.894 | 5.085 | 3.5 | ... | ... | ... | 5.3 | 19.27 | 8026.5 | 8042 |
| $\sigma$ | +0.04 | +2.8 | +0.002 | +0.5 | 0.006 | -0.012 | +2.2 | +0.1 | +1.6 | +0.3 | +1.9 | 0+0.06 | ... | ... |
| close | 6151.50 | 66.8 | -0.103 | 2.0 | 0.895 | 1.197 | 3.8 | 0.1 | 0.16 | 0.6 | ... | 5.3 | 19.26 | 8026.4 | 8042 |
| $\sigma$ | +0.04 | +3.6 | +0.006 | +0.5 | 0.006 | -0.003 | +2.0 | +0.1 | +1.0 | +0.4 | +2.4 | 0+0.07 | ... | ... |
| close + | 6151.50 | 66.7 | 0.104 | 2.1 | 1.005 | 5.088 | 2.9 | -0.5 | -0.3 | 0.6 | ... | 5.2 | 19.26 | 8027.5 | 8042 |
| $\sigma$ | +0.04 | +3.6 | +0.002 | +0.5 | +0.007 | -0.009 | +2.6 | +0.6 | +0.1 | +0.5 | +2.7 | 0+0.07 | ... | ... |
| wide | 6151.50 | 66.4 | -0.104 | 2.3 | 1.006 | 1.194 | 3.6 | 0.0 | -0.6 | 0.6 | ... | 5.3 | 19.26 | 8027.9 | 8042 |
| $\sigma$ | +0.04 | +4.0 | +0.007 | +0.3 | +0.006 | +0.006 | +2.0 | +0.1 | +1.1 | +0.5 | +2.5 | 0+0.09 | ... | ... |

Note. A superscript/subscript indicates the difference parameter’s 84/16 percentile from the best-fit value, respectively.
This relation comes from a private communication with them, which is restricted to FGK stars with $3900 \, \text{K} < T_{\text{eff}} < 7000 \, \text{K}$ and the accuracy of relation is better than 2% (Fukui et al. 2015). We measured the source color and brightnesses $(V - I)$, $S_2$, with the OGLE-IV instrumental magnitude from the light curve fitting. Note that because the source color and magnitudes are nearly identical for all models (see Table 2), we adopt values of parallax close+ model in the following analysis. We correct their extinction following the standard procedure by Yoo et al. (2004) using the red clump giants (RCG) as a standard candle. Figure 4 shows a color–magnitude diagram within the $2' \times 2'$ region around the source star with the OGLE-IV instrumental magnitude. The position of source and the measured RCG centroid $(V - I)_{\text{RC}} = (1.644, 15.27) \pm (0.011, 0.04)$ are shown as blue dot and red cross, respectively. Comparing the measured RCG centroid and the expected intrinsic position in this field $(V - I)_{\text{RC,0}} = (1.06, 14.38) \pm (0.07, 0.04)$ by Bensby et al. (2013) and Nataf et al. (2013), we obtain the intrinsic source color and magnitude as $(V - I)_{\text{S,0}} = (0.76, 18.40) \pm (0.07, 0.07)$ with the assumption that the source extinction are the same as that of the RCG. Note that the original reddening and extinction values in the standard magnitude in this field can be measured by the catalog of OGLE-III photometry map (Szymański et al. 2011) as $E(V - I) = 0.68 \pm 0.07$ and $A_I = 0.86 \pm 0.06$. We use these original values to obtain $A_H$ in Section 6.
intrinsic source color and magnitude to Equation (4), we obtain the angular source radius as \( \theta_s = 0.69 \pm 0.05 \) mas. From \( \theta_s, \rho \) and \( t_E \), we can calculate the angular Einstein radius \( \theta_E \) and the lens-source relative proper motion \( \mu_{\text{rel}} \).

\[
\theta_E = \frac{\theta_s}{\rho} > 0.22 \text{ mas}, \quad \mu_{\text{rel}} = \frac{\theta_E}{t_E} > 1.2 \text{ mas yr}^{-1}
\]

for the parallax close+ model. Table 3 shows the derived parameters for all degenerate models using each model’s values.

6. EXCESS BRIGHTNESSES FROM KECK AO IMAGES

We have a mass–distance relation via the microlens parallax \( \pi_E \). If we can measure the lens flux which gives us an additional mass–distance relation, we could measure the mass and distance of the lens uniquely. We conducted high angular resolution observations using adaptive optics in order to measure the lens flux, excluding as much flux from unrelated stars as possible.

6.1. Observations and the Photometry

We observed OGLE-2012-BLG-0950 with the NIRC2 instrument on Keck II on 2013 July 18. We used the Wide camera with a pixel scale of 0.04 arcsec. We took 15 dithered \( H \) frames with an exposure time of 30 s. We performed dark and flatfield corrections in the standard way. Furthermore, OGLE-2012-BLG-0950 was observed as part of the VVV survey (Minniti et al. 2010) using the VISTA 4 m telescope at ESO. We extracted 3 arcmin VVV \( JHK \) images centered on the target. We used the suite of tools developed as part of astrOmatic (Bertin et al. 2002). We analyzed the PSF of the images using PSFEx, then we measured the fluxes with SEextactor (Bertin & Arnouts 1996) using these PSF models. We cross identified 2MASS stars with VVV sources, and finally derived the calibration constant for Keck \( H \) band photometry.

Figure 3 shows the field as observed by VVV and by Keck. First, we notice that the source in VVV is resolved in 2 objects.
with Keck. Using the astrometry on the amplified source, we are able to identify the source-lens of the microlensing as being the star that is marked on the frame. Its coordinates are (R.A., Decl.)J2000 = (18:08:04.620, −29:43:53.43). It has a Keck H band magnitude of

\[ H_{\text{target}} = 16.89 \pm 0.02 \]  

in the 2MASS magnitude system. The blend to the south is at coordinates (R.A., Decl.)J2000 = (18:08:04.612, −29:43:53.48) and is slightly fainter at \( H = 16.99 \pm 0.03 \).

### 6.2. The Excess Flux

Considering a full width at half maximum of the target FWHM\(_{\text{target}}\) = 90 mas in Keck image and the lens-source relative proper motion, \( H_{\text{target}} \) includes the lens flux plus the source flux. Here, we derive the excess brightness by subtracting the source system brightness, which can be used as the lens flux or its upper limit. Hereafter, we represent our derived values for only the parallax close+ model as the parallax model, unless otherwise stated.

Unfortunately, we do not have a light curve in the \( H \)-band, so we derived the source \( H \) magnitude as \( H_{S,0} = 17.55 \pm 0.12 \) by converting the magnitude from \( (V - I, H)_{S,0} \) with the color–color relation by Bessell & Brett (1988). Next we applied the \( E(V - I) \) and \( A_V \) values derived in Section 5 and estimated the extinction in \( H \)-band as \( A_H = 0.25 \pm 0.02 \) (Cardelli et al. 1989). The magnification at the time of Keck observation at HJD’ = 6491.88 is \( A = 1.005 \) for the parallax model. Thus, the source apparent \( H \) magnitude at the time is \( H_{S,\text{Keck}} = 17.78 \pm 0.12 \). This value is converted to the 2MASS magnitude system from the Bessell & Brett system using Equations (12)–(15) in Carpenter (2001). Subtracting this
from $H_{\text{target}}$ of Equation (5), we derive the excess brightness of

$$H_{\text{excess}} = 17.52 \pm 0.10.$$  \hspace{1cm} (6)

6.3. Probability of the Contamination Fraction $f$

We next consider the probability that part of the $H_{\text{excess}}$ come from stars other than the lens. To estimate the probability of the contamination fraction $f$, we consider the following three possible sources of contamination: unrelated ambient stars, a companion to the source star, and a companion to the lens star (Batista et al. 2014; Fukui et al. 2015).

6.3.1. Unrelated Ambient Stars

First, we estimate the probability of contamination owing to unrelated ambient stars. Counting the number of stars with $H > H_{\text{excess}}$ on the Keck image, we estimate the number density of stars with brightnesses corresponding to $0 < f < 1$ as 0.68 arcsec$^{-2}$. Similarly, the number density of stars corresponding to $0.5 < f < 1$ and $0 < f < 1$ are estimated as 0.11 arcsec$^{-2}$ and 0.01 arcsec$^{-2}$, respectively.

We can resolve an ambient star only if it is separated from the source by FWHM = 90 mas or more. Thus, the probability of contamination owing to unrelated ambient stars within 90 mas around the source, $P_{\text{amb}}$, is $P_{\text{amb}}(0 < f < 1) = 1.74\%$, $P_{\text{amb}}(0.5 < f < 1) = 0.29\%$ and $P_{\text{amb}}(0.9 < f < 1) = 0.03\%$.

6.3.2. Companion to the Source Star

Second, we estimate the probability of contamination owing to a companion to the source. Because we can detect a companion on the Keck images if the companion is located far enough from the source and the light curve will be affected if the companion is located close enough to the source, an undetectable companion should be located between the two detection limits. We put the distant limit as 90 mas from the FWHM of the target. For the close limit, we consider $a_{\text{SC,low}}$ as defined in the following.

As we can find in the bottom inset in Figure 3, the OGLE $I$ data in the light curve is sensitive enough to a deviation with an amplitude of $\Delta A \approx 0.1$ and a duration of hundreds of days. That means we can detect a binary source signal when a companion to the source is magnified with an amplitude of $\geq 0.1f_5$ and a duration of hundreds of days. Such small signals cannot be detected if they are longer than $10f_5 \sim 670$ days. We assume that we can detect the signal caused by a magnified companion to the source star when the time variation of the companion flux is larger than $0.1f_5$ within $10f_5$, that is, it requires

$$A(u_{C,0},f_C) - A(\sqrt{5^2 + u_C^2},f_C) \leq 0.1f_5$$  \hspace{1cm} (7)

to be “undetected”, where 5 comes from half of “10” $f_5$, and $u_{C,0}$ and $f_C$ are the impact parameter and the flux in the $I$-band of the companion star, respectively. $A(u_C)$ is the magnification when the companion is located at $u_C$ and we use the magnification owing to a single lens $A(u) = (u^2 + 2)/u \sqrt{u^2 + 4}$ as $A(u_C)$. Because $A(\sqrt{5^2 + u_C^2}) < A(5)$, we can express the condition more conservatively,

$$A(u_{C,0},f_C) - A(5)f_C \leq 0.1f_5$$

$$u_{C,0} \geq u_C \left(1.00275 + 0.1\frac{f_5}{f_C}\right) \equiv u_{C,0,\text{low}}.$$  \hspace{1cm} (8)

The $u_{C,0,\text{low}}$ is a lower limit of the impact parameter of the companion to be undetected and it gives us a lower limit of the separation between the source and the companion as $s_{\text{SC}} > s_{\text{SC,low}}$. When the time when the companion is located at $u_C = u_{C,0}$ is defined as $t_{C,0}$, it is possible to give the most conservative lower limit as $s_{\text{SC}} > s_{\text{SC,low}} \equiv |u_{C,0,\text{low}} - u_0|$ when $t_{C,0} = t_0$ and the closest companion crosses the same side as the source relative to the lens. Thus, we derive the lower limit of the semimajor axis of the source system as $a_{\text{SC,low}} \equiv D_8\theta_{E,\text{SC,low}}$.

To calculate $u_{C,0,\text{low}}$ in $a_{\text{SC,low}}$, we obtain $M_{t,C}$ by convolving

$$M_{t,C} = H_{\text{excess}} - 2.5\log f - 5\log(D_8/10\text{pc}) \rightarrow M_{t,C}$$

using PARSEC isochrones version 1.2S (Bressan et al. 2012; Chen et al. 2014, 2015; Tang et al. 2014) with $D_S = 8$ kpc for the calculation of $f_5/f_C = 10^{-0.4(M_{t,C} - M_1)}$ in Equation (8) where $M_{t,S}$ and $M_{t,C}$ are absolute $I$ magnitudes of the source and the companion.

Because $a_{\text{SC,low}}$ and $\theta_{E}$ vary with $f$, we calculate the probability of contamination owing to a companion to the source using the following formula (Fukui et al. 2015, cf.),

$$P(f_5 < f < f_2) = F_{\text{binary}} \times \int_{f_5}^{f_2} F_{\text{in}}(f) \times F_{\text{sc}}(f)df.$$  \hspace{1cm} (9)
where $F_{\text{binary}}$ is the multiplicity of FGK-dwarfs, $F_{\text{sc}}(f)$ is the fraction of binaries with a separation of $a_{\text{SC,lim}}(f) < a_c < 90 \text{ mas} \times D_\text{s}$ and $F_{\text{sc}}(f)$ is the fraction of binaries with a mass ratio between $q_\text{f}(f)$ and $q_\text{f}(f + df)$. We derive $M_\text{f}(f)$ from $M_{\text{H},C}(f)$ using the PARSEC isochrones version 1.2S for the calculation of $q_\text{f}(f)$. We use the distribution of parameters for FGK binaries by Raghavan et al. (2010): $F_{\text{binary}} = 0.46$ as the multiplicity, a log normal distribution with the mean of $\log P_\text{f}(\text{days}) = 5.03$ and the standard deviation of $\sigma_{\log P_\text{f}} = 2.28$ as the period distribution and Figure 16 in the paper as the mass ratio distribution. We apply the period distribution by converting $a_\text{c}$ to a period using Kepler’s 3rd law. Then we obtain $P_{\text{sc}}(0 < f < 1) = 19.3\%$, $P_{\text{sc}}(0.5 < f < 1) = 2.2\%$, $P_{\text{sc}}(0.9 < f < 1) = 0.5\%$ for the parallax close+ model.

6.3.3. Companion to the Lens Star

Next, we estimate the probability for contamination owing to a companion to the lens star. In this case, we again use Equation (9) to calculate the probability because the lens mass, the distance to the lens and the companion mass values are dependent on the $f$ value in addition to $\theta_\text{E}$. However, the definition of $F_{\text{sc}}(f)$ is different here. To place close limits on $a_\text{c}$, we consider the shear given by a hypothetical lens companion (Batista et al. 2014). Assuming that we can detect the shear effect when the width of the central caustic created by the companion, $w_\text{c} \approx 4q_\text{c}/s_\text{c}^2$, is larger than the width of that by the planet, $w$, we can place a detection limit as $w_\text{c} < w$, where $s_\text{c}$ is the projected separation normalized by $\theta_\text{E}$ between the host and a hypothetical companion of the lens. This inequality gives us a close limit on $a_\text{c}$.

Adopting 90 mas as the distant limit again, $F_{\text{sc}}(f)$ here is defined as the fraction of binaries with a projected separation of $2\sqrt{q_\text{c}(f)/w} \theta_\text{E}(f) D_\text{s}(f) < a_c < 90 \text{mas} \times D_\text{s}(f)$ where $w = 4q_\text{c}/(s - s_\text{c})^2 \approx 0.016 \sim 0.02$ for the close models. We apply $w = 0.02$ to the wide model as well as the close model considering that the shear effect by a hypothetical companion is almost equal to that for the close model because the magnification map on the source trajectory for the wide models is almost the same as the close models. Note that uncertainties arising from any assumptions here do not affect the results, largely because they are too small compared to the range of $a_c$ values we are considering.

We calculate $q_\text{f}(f)$ for each $f$ value using isochrone models as well as the source companion case. Because a primary of the lens can be either an M dwarf or FGK dwarf depending on the $f$ value, we use the distributions for FGK-dwarf binaries by Raghavan et al. (2010) for $M_{\text{prim},f}(f) > 0.7$ and that for KM-dwarf binaries by Ward-Duong et al. (2015) for $M_{\text{prim},f}(f) \leq 0.7$. Ward-Duong et al. (2015) gives $F_{\text{binary}} = 0.347$ as the multiplicity, a log normal distribution with the mean of $\log a_\text{c,(au)} = 0.77$ and the standard deviation of $\sigma_{\log a_\text{c}} = 1.34$ as the distribution of the projected separation and a mass ratio distribution that is flat for $0.2 < q_\text{c} < 1$ and 0 for $q_\text{c} < 0.2$. Integrating Equation (9) with these distributions and the detection limits, we derive $P_{\text{sc}}(0 < f < 1) = 13.7\%$, $P_{\text{sc}}(0.5 < f < 1) = 7.2\%$ and $P_{\text{sc}}(0.9 < f < 1) = 3.9\%$ for the parallax close+ model.

Figure 6. Mass–distance relations for the parallax close+ model. The relations from $M_{\text{H,L}}$ are shown in red, blue, and green for a contamination fraction of $f = 0, 0.5$, and 0.9, respectively. The dashed lines indicate 1σ error including the uncertainty of the distance to the source, the lens age, and the lens metallicity in addition to the uncertainty of our measuring. The black lines are the mass–distance relation from $\pi_\text{E}$.

6.3.4. Total Probability of Contamination

Finally, by summing the probabilities for the three sources mentioned above, we can calculate the following total probability of contamination: $P(0 < f < 1) = 37.4\%$, $P(0.5 < f < 1) = 10.7\%$, $P(0.9 < f < 1) = 4.6\%$ assuming the parallax close+ model. Therefore, the probabilities of the contamination fraction not exceeding $f$ shown in Table 3 can be calculated as $P(f = 0) = 1 - P(0 < f < 1) = 65.3\%$, $P(f \leq 0.5) = 1 - P(0.5 < f < 1) = 90.3\%$ and $P(f \leq 0.9) = 1 - P(0.9 < f < 1) = 95.6\%$ for the parallax close+ model. The probabilities for all parallax models are shown in Table 3.

7. LENS PROPERTIES

In this section, we constrain the lens properties from microlens parallax and the lens flux. All results are summarized in Table 3.

We have two mass–distance relations for the parallax model. One is the lens absolute magnitude,

$$M_{H,L} = H_L - A_{H,L} - 5 \log \frac{D_L}{10 \text{pc}}$$

$$\left(H_L = H_{\text{excess}} - 2.5 \log (1 - f)\right)$$

where $f$ is the fraction of contamination flux to excess flux and $A_{H,L}$ is an extinction for the lens located at $D_\text{L}$ and we adopted $A_{H,L} = A_H D_\text{L}/D_\text{S}$ following Fukui et al. (2015). This is converted to a relationship between the host mass $M_{\text{host}}$ and the distance $D_\text{L}$ by using isochrone models from the PARSEC isochrones, version 1.2S (Bressan et al. 2012; Chen et al. 2014; Tang et al. 2014; Chen et al. 2015). The other mass–distance relation is from the microlens parallax $\pi_\text{E}$:

$$M_{\text{host}} = \frac{1}{1 + q \kappa \pi_\text{rel}^2}$$

where $\kappa = 8.144 \text{ mas} M_\odot^{-1}$ and $\pi_\text{rel} = \text{au} (1/D_L - 1/D_S)$. 

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Figure 6 shows these two relations for the parallax close+ model. The black lines are the mass–distance relation from $\theta_E$. The red, blue, and green lines indicate the relation from $M_{\text{lu}},$ for the case of $f = 0, 0.5,$ and $0.9,$ respectively. Dashed lines indicate $1\sigma$ errors and they include the uncertainty of the distance to the source of $D_S = 8.0 \pm 1.6$ kpc, the lens age of $<13$ Gyr, and the lens metallicity of $[\text{Fe}/\text{H}] = -0.05 \pm 0.20$ in addition to the uncertainty of our measurements. We adopt the metallicity distribution of nearby M- or late K-dwarf stars (Gaidos et al. 2014, e.g.) for the metallicity. Note that the dependency on age is much weaker than that on metallicity in the region of the parallax solution.

The region overlapping these two relations corresponds to the allowed solution. For $f = 0$, the host mass is $M_{\text{host}} = 0.57^{+0.06}_{-0.10} M_{\odot}$, and the distance is $D_L = 2.6^{+0.5}_{-0.7}$ kpc, and the planet mass is $M_p = 35^{+10}_{-6} M_{\oplus}$, its projected separation is $r_p = 2.6^{+0.3}_{-0.5}$ au, and the three-dimensional star–planet separation is statistically estimated as $a = 3.1^{+1.5}_{-0.5}$ au with a circular orbit assumption (Gould & Loeb 1992). From this solution, we can calculate the angular Einstein radius and the relative lens-source proper motion as $\theta_E = 1.09^{+0.16}_{-0.10}$ mas and $\mu_{\text{rel}} = 5.8^{+0.6}_{-0.8}$ mas yr$^{-1}$, respectively. The solution of $M_{\text{host}} = 0.50^{+0.05}_{-0.09} M_{\odot}$, $D_L = 2.9^{+0.0}_{-0.8}$ kpc for $f = 0.5$ is consistent with that for $f = 0$ within 1$\sigma$. In addition, the solution for $f = 0.9$ is $M_{\text{host}} = 0.33^{+0.02}_{-0.06} M_{\odot}$, $D_L = 3.6^{+0.5}_{-0.8}$ kpc. Our estimate for these contamination probabilities is discussed in Section 6.3 and summarized also in Table 3.

In any case, the host star is an M/K dwarf and the planet is a Neptune/sub-Saturn mass planet. All solutions of the other degenerate parallax models are similar to these results as shown in Table 3. We present the mean value of the 8 parallax solutions with $f = 0$ and $f = 0.5$ without any weight as “Mean” in Table 3. Here, the contributions of solutions with $f > 0.5$ are negligible.

8. DISCUSSION AND CONCLUSION

We analyzed the microlensing event OGLE-2012-BLG-0950. A negative perturbation in the microlensing light curve consistent with a low-mass planet was detected (Abe et al. 2013). All the models we analyzed have a planetary mass ratio, $q \simeq 2 \times 10^{-4}$. We could not detect a significant finite source effect because the source did not cross any caustic, but we did detect a parallax signal. The parallax solutions indicate a Neptune/sub-Saturn mass planet with mass of $M_p = 35^{+10}_{-6} M_{\oplus}$ around an M/K-dwarf host with mass of $M_{\text{host}} = 0.56^{+0.12}_{-0.10} M_{\odot}$. We measured the lens mass by combining microlens parallax and the lens flux obtained by Keck AO observations. This is the first case in which the lens mass was measured using only microlens parallax and the lens flux.

The planet orbits outside of the snow line of the host star and has a mass between that of Neptune and Saturn, $M_p = 35^{+10}_{-6} M_{\oplus}$. Planets with this mass range (intermediate mass, hereafter) are predicted to be rare inside the snow line, but to be common like Neptune- or Saturn- mass planets outside the snow line according to the core accretion theory (Ida & Lin 2004; Ida et al. 2013). A paucity of intermediate mass planets orbiting close to their metal-poor host stars is confirmed (Beaugé & Nesvorný 2013). On the other hand, the predicted relative abundance outside the snow line has not been confirmed yet. Figure 7 shows the distribution of the exoplanets discovered so far. The solution of our parallax model is indicated as the purple filled circle located just around the valley of the bimodal mass distribution histogram on the left side of the figure. Note that this distribution is not corrected for detection efficiency. Only a few intermediate mass planets orbiting outside the snow line have been discovered by the RV and microlensing methods. The parallax model of this work could be the second such intermediate mass exoplanet with mass measurement, following OGLE-2012-BLG-0026Lb (Han et al. 2013; Beaulieu et al. 2016). The mass of other intermediate mass planets are estimated by Bayesian analysis (Miyake et al. 2011; Poleski et al. 2014; Skowron et al. 2016b). However, the Bayesian estimates depend on the choice of prior (Bennett et al. 2014; Skowron et al. 2015).

In a future space-based microlensing survey by WFIRST (Spergel et al. 2015) or Euclid (Penny et al. 2013), and in the survey of Campaign 9 of the K2 Mission (Henderson et al. 2015) conducted from 2016 April to 2016 July, or the Spitzer microlensing campaign from 2014 (Yee et al. 2015), it is important and easy to determine the lens mass for each event by combining microlens parallax and lens flux as pointed out by Yee (2015) for the following reasons. First, space- and ground-based simultaneous observations are expected to obtain microlens parallax for a significant fraction of events regardless of the number of the lens bodies, in contrast to the finite source effect which can be obtained only by observing the peak of the high-mag event or caustic crossing. Second, for low-mass and nearby lenses, the mass–distance relations derived from flux and from $\theta_E$ are partially degenerate (see Figure 2 in Yee 2015 or Figure 7 in Fukui et al. 2015) although we can obtain $\theta_E$ by the measurement of astrometric microlensing effects with the precision of WFIRST (Gould & Yee 2014). Third, the cases without detection of $\theta_E$ like this event are expected to increase, even for planetary events, because a higher precision and higher cadence survey can detect more subtle planetary signals, including cases without crossing caustics (Zhu et al. 2014). Finally, it is possible to measure the lens fluxes even after the events by follow-up observations with high angular resolution, and ultimately, WFIRST and Euclid can routinely measure the

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lens fluxes as part of the survey observations. Our analysis for the parallax model is the first demonstration of the mass measurement from only microlens parallax and the lens flux, and thus it has particular significance for the developing era of space-based microlensing.

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APPENDIX

XALLARAP ANALYSIS

If the source star is in a binary system, the orbital motion of the source star can also measurably affect the trajectory of the source during a microlensing event. This effect, called xallarap, requires additional parameters which define orbital elements of the source system, whereas we know the Earth’s orbital elements for the parallax effect. This model requires 7 additional fitting parameters to the standard binary model, the direction toward the Earth relative to the source orbital plane, RA, Decl, the orbital period P2, the orbital eccentricity e, and the perihelion time tper, in addition to ξE = (ξE, ξE, E) which is analogous to π for microlens parallax. We omitted e and tper as fitting parameters by assuming a circular orbit. Here, we describe the results of our xallarap analysis and comparison with the parallax model, and discuss the possibility.

A.1. Constraint by the Companion Mass Upper Limit

Kepler’s 3rd law gives us a relation of the source orbit;

\[
\xi_E = \frac{\text{au}}{D_\phi \theta_s / \rho} \left( \frac{M_C}{M_s} \left( \frac{M_s + M_C}{M_s + M_C \text{ 1yr}} \right) \right)^{2/3}
\]

(12)

where \(M_s\) and \(M_C\) are the masses of the source and its companion, respectively. The \(\xi, \rho, \text{ and } P_2\) are fitting parameters, \(\theta_s\) is measured, and \(D_\phi\) and \(M_s\) are reasonably constrained by the Galactic model combined with the source color and magnitude values. Then, for a given MCMC chain, \(M_C\) can be calculated from this relation.

In xallarap fitting with no constraints, we found that the light curve prefers a solution with an unrealistically massive source companion with a mass of \(M_C \sim 400 M_\odot\). Thus, we conducted the fitting with the following constraint. We can place an upper limit of the companion mass \(M_C < M_{C\text{,max}}\) from Equation (6) of the excess flux with an assumption that the companion is not a stellar remnant. That corresponds to placing an upper limit of \(\xi_E\) as \(\xi_{E, \text{max}}\) where \(\xi_{E, \text{max}}\) is defined as \(\xi_{E, \text{max}}\) of Equation (12) with \(M_C = M_{C\text{,max}}\) (Bennett et al. 2008; Sumi et al. 2010). We applied the additional \(\chi^2\) penalty presented by Bennett et al. (2008) to each link of MCMC fitting;

\[
\chi^2_{\text{orb}} = \Theta(\xi_E - \xi_{E, \text{max}}) \left( \frac{\xi_E - \xi_{E, \text{max}}}{\sigma_{\xi_{E, \text{max}}}} \right)^2
\]

(13)

where \(\Theta\) is the Heaviside step function and we applied 7% to \(\sigma_{\xi_{E, \text{max}}}\), the uncertainty of \(\xi_{E, \text{max}}\) with the consideration of uncertainty on \(\theta_s\) as given in Section 5.

We derive \(D_\phi = 8.0 \pm 1.6\) kpc by a Bayesian analysis using the Galactic model (Han & Gould 2003) as the prior distribution constrained by the observed \(E_\phi\) value, \(M_\odot = 1.02 \pm 0.12 M_\odot\) and \(M_{C, \text{max}} = 1.10 \pm 0.19 M_\odot\) from the color and brightness of the source and blending. To calculate \(\xi_{E, \text{max}}\), we use the lower or upper limit value for each parameter so that it makes \(\xi_{E, \text{max}}\) larger for a conservative constraint, namely, we use \(D_\phi = 6.4\) kpc, \(M_\odot = 0.9 M_\odot\), \(M_{C, \text{max}} = 1.2 M_\odot\) for the calculation of \(\xi_{E, \text{max}}\). Note that we adopt \(1.2 M_\odot\) as \(M_{C, \text{max}}\) considering that there are very few \(M > 1.2 M_\odot\) stars in our galaxy (Gould 2000b; Bensby et al. 2013). For \(\rho\) and \(P_2\), we used each link’s values to calculate the \(\xi_{E, \text{max}}\).

The “xallarap” models in Table 2 are our results of xallarap fitting with a circular orbit and the \(\xi_{E, \text{max}}\) constraint above. We find smaller \(\chi^2\) values than that from the parallax models by \(\Delta \chi^2 > 27\). Note that including eccentricity intends to fit systematics in the baseline and does not improve a model significantly, so we do not consider eccentric orbits according to Occam’s razor.

A.2. Constraint on \(\rho\) and Lens Properties

Xallarap models place a lower limit on \(\rho\), whereas the parallax models do not. This is because the \(\xi_{E, \text{max}}\) constraint is equivalent to placing a lower limit of \(\rho\) as

\[
\xi_E < \xi_{E, \text{max}} \Leftrightarrow \rho > \rho_{\text{min}}
\]

\[
= \frac{\theta_s D_\phi \xi_E}{M_{C, \text{max}} (M_{C, \text{max}} + M_\odot)^{-2/3} \rho_{\xi}^{2/3}}.
\]

(14)

Combining it with the upper limit from the finite source effect, we can constrain the \(\rho\) value with \(\sim 30\%\) uncertainty. This is the first case of \(\rho\) being constrained with neither a significant finite source effect nor a parallax effect. Then, we can calculate a \(\xi_E\) value from \(\rho\) and \(\theta_s\),

\[
\theta_E = 0.20^{+0.04}_{-0.08} \text{ mas}, \quad \mu_{\text{rel}} = 1.1^{+0.2}_{-0.4} \text{ mas yr}^{-1}
\]

for the xallarap close+ model. Table 3 shows these values for all degenerate models using each model’s values.

In principle, we can determine a mass and distance of the lens star by combining the \(\theta_E\) and \(H_T\), the lens flux in \(H\)-band extracted from Keck AO observations (Batista et al. 2014; Fukui et al. 2015, e.g.,) because \(\theta_E\) gives us a mass–distance relationship;

\[
M_{\text{host}} = \frac{1}{1 + q \kappa 
\]

(15)

and we can convert the lens flux into another mass–distance relationship using a mass–luminosity relation. However, we encounter a problem with this. As shown in Table 4, the uncertainty of the mass of the companion to the source star in the xallarap models is very large and the upper limit of \(\sim 1.8 M_\odot\) is larger than \(M_{C, \text{max}} = 1.2 M_\odot\), the maximum mass of the
especially, it is more sensitive to the uncertainties $M$. Then, we can place a lower limit of $6.4 \text{kpc}$ and the other parameters, combination of $M_{\text{C}}$ and $M_{\text{S}}$, exceeds its limit, which is slightly magnified. The red arrow indicates that the red dashed line is just an upper limit. The black lines are the mass–distance relation arising from $\theta_E$.

A.3. Comparison with Parallax Model

Xallarap models have $\chi^2$ differences from the parallax models by $\Delta \chi^2 > 27$ for 3 dof difference. Poindexter et al. (2005) analyzed 22 events where a parallax model improves their light curve fits compared to the standard model. According to their analysis, there are 3 events that prefer a xallarap model to a parallax model by $\Delta \chi^2 > 25$ in all events they analyzed. They regard it as a strong indication that the light curves of the 3 events have been distorted by xallarap. Here, we investigate whether the xallarap signal is real or unreal in the analyses below.

We first plot $\chi^2$ values of the best-fit xallarap model at a fixed $P_z$ value within $100 \leq P_z$ (days) $\leq 1500$ in Figure 9. They are shown as $\Delta \chi^2$ values compared to the parallax close+ model. One of the standard ways to exclude a xallarap scenario is to show that just a narrow range of $P_z$ indicates a (small) preference in favor of the xallarap model over the parallax model, and then showing the very small probability that a binary system whose period is in such a narrow range happens to be microlensed (Bennett et al. 2008). For this event, however, the favored $P_z$ region compared to the parallax model is very broad, even with the $\chi^2_{\text{orb}}$ constraint. Therefore, we cannot exclude the xallarap scenario by this approach.

Next, we investigate where the signal comes from in the light curve. Figure 10 shows the same one as Figure 3, but shows the difference between the xallarap close+ model and the parallax close+ model. The $\chi^2$ difference comes from the data in 2013 which is slightly magnified. However, the preference to the parallax model comes almost exclusively from MOA data and

| Model       | $M_C$ (M$_\odot$) | $H_C$ (mag) |
|-------------|------------------|-------------|
| Xallarap    |                  |             |
| close+      | 0.57–1.71        | 16.3–20.4   |
| close−      | 0.56–1.71        | 16.3–20.5   |
| wide+       | 0.62–1.81        | 16.1–20.7   |
| wide−       | 0.60–1.76        | 16.2–20.2   |

Note. The range of 1σ error is shown for each parameter.
the preference from OGLE data is only \( \Delta \chi^2 \sim 2 \). The MOA data is more easily affected by systematics especially in a low-magnification part because the average seeing on the MOA site is worse than that on the OGLE site. In addition to this inconsistency between MOA and OGLE, as shown in Table 4, a significant fraction of MCMC chains of every xallarap model indicates unphysically large masses of the companion even with the mass constraint imposed. These two facts are strong evidence against the xallarap model as the true model. Thus we conclude that the signal modeled as xallarap in this model is unlikely to be owing to a real xallarap effect. Instead it is most likely to arise from low-level systematics in the MOA data.

Finally, we fit the data consisting of only the anomaly part of MOA data (39 points) and entire OGLE I data (1273 points) with both parallax and xallarap with no constraint models because the results are expected to be less affected by systematics. Note that we use a potion of MOA data so that the best fit will not change largely except \( \tilde{\xi}_E \) or \( \tilde{\xi}_c \). The xallarap model improves the fit by \( \Delta \chi^2 = 6.5 \) compared to the parallax model for 3 dof difference. This indicates that the xallarap model is preferred by only \(< 2\sigma\). Moreover, we compared these two models by the Bayesian information criterion (BIC), which is another common statistical criteria including a penalty term for the number of fitting parameters, \( \text{BIC} = \chi^2 + n_{\text{param}} \ln(N_{\text{data}}) \) (Burnham & Anderson 2002). The \( \Delta \text{BIC} \) is \( \text{BIC}_{\text{parall}} - \text{BIC}_{\text{xallarap}} \sim -15 \) and this criterion prefers the parallax model rather than the xallarap model.

Considering these facts, we conclude that the parallax model is preferred over the xallarap model. Thus, we dealt with only the parallax scenario in the main part of the paper.

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Figure 10. As for Figure 3, but showing the difference between the parallax close+ model and the xallarap close+ model. A positive \( \Delta \chi^2 \) value indicates a smaller \( \chi^2 \) of the xallarap close+ model.
