Efficacy of pregabalin in post-traumatic peripheral neuropathic pain: a randomized, double-blind, placebo-controlled phase 3 trial

John Markman1 · Malca Resnick2 · Scott Greenberg2 · Nathaniel Katz3 · Ruoyong Yang2 · Joseph Scavone2 · Ed Whalen2 · Gabriela Gregorian2 · Bruce Parsons2 · Lloyd Knapp2

Received: 15 April 2018 / Revised: 17 August 2018 / Accepted: 16 September 2018 / Published online: 21 September 2018 © The Author(s) 2018

Abstract

The growing need for symptomatic treatment of post-traumatic neuropathic pain (PTNP) continues to be unmet. Studies evaluating the efficacy of pregabalin for reducing neuropathic pain following trauma and surgery yielded positive results over ≤ 8-week treatment. To assess the efficacy and tolerability of pregabalin over 3 months in patients with PTNP, a randomized, double-blind, placebo-controlled, parallel-group trial evaluated patients with PTNP at 101 centers in 11 countries—the longest, largest such trial. Adults diagnosed with PTNP were randomly assigned (1:1) to 15 weeks of pregabalin (flexibly dosed 150–600 mg/day) or matching placebo. Primary efficacy analysis was by mixed-model repeated measures comparing change from baseline to week 15 in weekly mean pain scores between active and placebo groups. Evaluable patients included 274 in the pregabalin group and 265 in the placebo group. Trauma was surgical in 49.6% of patients, non-surgical in the remainder. The primary efficacy analysis showed no statistically significant difference between pregabalin and placebo groups in the change from baseline to week 15 [mean difference, −0.22 points (95% confidence interval, 0.54–0.10); \( p = 0.1823 \)]. However, comparisons for key secondary outcome measures yielded \( p \) values < 0.05 favoring pregabalin. Consistent with the known safety profile of pregabalin, the most common adverse events were dizziness and somnolence (14.6 and 9.9% of patients, respectively) with pregabalin (vs 4.2 and 3.4% with placebo). These findings demonstrate the feasibility of conducting a large, phase 3 registration trial in the heterogeneous PTNP study population.

ClinicalTrials.gov NCT01701362.

Keywords Neuropathic pain · Post-traumatic neuropathic pain · Post-surgical neuropathic pain · Pregabalin

Introduction

Diverse types of nerve injury are recognized as triggers of chronic post-traumatic neuropathic pain (PTNP), including post-surgical syndromes [1–3]. Tissue injury may chronically alter peripheral nociceptive processing, shifting pain from acute to chronic [4, 5]. The need for symptomatic treatment of PTNP is increasing but remains inadequately addressed [6–8].

Pregabalin, an alpha2-delta (\( \alpha_2\delta \)) ligand (gabapentinoid), is approved in the United States for the treatment of three neuropathic pain (NeP) conditions: diabetic peripheral neuropathy (DPN), post-herpetic neuralgia (PHN), and post-spinal cord injury (SCI) [9]. An 8-week randomized clinical trial demonstrated the efficacy of pregabalin for the management of chronic post-traumatic/post-surgical pain [10].

A study of longer duration was designed to meet the US regulatory standard for a chronic pain indication: 12 weeks of maintenance or fixed dosing [11]. Methodologic features were incorporated to increase assay sensitivity for the detection of an analgesic signal in this heterogeneous patient population [12]. The primary objective was to compare the efficacy of pregabalin (flexibly dosed, 150–600 mg/day) versus placebo in the treatment of PTNP. Secondary objectives

Electronic supplementary material The online version of this article (https://doi.org/10.1007/s00415-018-9063-9) contains supplementary material, which is available to authorized users.

* John Markman
  John_Markman@urmc.rochester.edu

1 Translational Pain Research Program, Department of Neurosurgery, University of Rochester Medical Center, 601 Elmwood Avenue, Box 670, Rochester, NY 14642, USA
2 Pfizer Inc, New York, NY, USA
3 Analgesic Solutions, Natick, MA, USA
compared the efficacy of pregabalin vs placebo with respect to overall status, pain-related activity limitation, and sleep, in addition to safety and tolerability assessments.

Methods

Study patients

Eligible patients were aged ≥ 18 years and had PTNP for ≥ 6 months after a surgical or non-surgical traumatic event (e.g., history of a motor vehicle accident, fall, sports injury, knee or hip replacement, hernia repair, thoracotomy, mastectomy, focal/localized burns, or crush injury), a mean score of ≥ 4 in pain recall for the past week at screening, and a mean score of ≥ 4 and ≤ 9 on a 0–10 numeric rating scale (NRS) of average pain (0, “no pain” to 10, “worst possible pain”) based on ≥ 4 daily diary scores from the last week of a single-blind baseline screening period (5–14 days before randomization). Peripheral nerve(s) implicated in the pain was identified to confirm nerve trauma, and pain was categorized as neuropathic based on prespecified criteria (i.e., neurologic exam, study-specific PTNP assessment including use of the PainDETECT questionnaire to identify neuropathic components of back pain [13]) and diagnostic tests (e.g., electromyography, nerve conduction tests, skin or nerve biopsy) if available. While PainDETECT was a screening assessment and used as part of the initial diagnostic assessment, since it is not specifically validated for this indication, this instrument itself did not determine eligibility in the study.

Each neuropathic symptom or sign, mapped separately, was submitted to a team of independent neurologists (contracted by Analgesic Solutions, Natick, MA, USA) who determined the plausibility of matching a PTNP syndrome with respect to the history, anatomic distribution of reported pain, and associated signs identified on neurologic examination in the corresponding body region.

Exclusion criteria included NeP due to PHN, DPN, complex regional pain syndrome, and other conditions; other sources of pain that might confound assessment of PTNP; disallowed concomitant medications; nonpharmacologic treatments for PTNP; severe or acute medical or psychiatric conditions; or clinically significant laboratory abnormalities. Patients scoring ≥ 15 on the Patient Health Questionnaire (PHQ-8) at screening or who were at risk based on Columbia-Suicide Severity Rating (C-SSRS) responses were recommended for evaluation by a mental health professional prior to randomization [14–16]. Prohibited medications included opioids, local anesthetics, topical and intraspinal steroids, antiepileptics, and antidepressants. Allowed medications included stable regimens of nonsteroidal anti-inflammatory drugs (NSAIDs), non-opioid analgesics, antidepressants [including serotonin-specific reuptake inhibitors (SSRIs), tricyclic antidepressants, and serotonin–norepinephrine reuptake inhibitors (SNRIs)], tramadol and triptans, and/or sleep medications; acetaminophen ≤ 3 g/day was allowed as rescue medication.

Study design and procedures

Eligible patients were randomized at 101 centers in 11 countries (Bulgaria, Canada, Denmark, Germany, Hungary, Poland, Romania, Sweden, South Africa, South Korea, and the United States). Following a single-blind screening period, the 15-week double-blind treatment period comprised 3 weeks of dose titration/optimization and 12 weeks of maintenance treatment (Supplementary Fig. S1). After randomization, clinic visits occurred every 3 weeks.

Patients were randomized (1:1) to pregabalin or matching placebo. The pregabalin dose was individually optimized via telephone contact to 150, 300, 450, or 600 mg/day over the titration period, with 4 days at each dose before titration to the next level. Dose adjustments were not allowed during the maintenance period, except for a single dose reduction if the investigator judged it necessary for tolerability.

Outcome measures

The primary outcome was pain rated in a diary completed by telephone each evening between 7 pm and midnight. Patients were asked to select the number that best described their NeP during the past 24 h on an 11-point NRS ranging from 0 (“no pain”) to 10 (“worst possible pain”). During this call, patients were also asked to select the number that best described how NeP had interfered with their sleep during the past 24 h, on an NRS from 0 (“pain does not interfere with sleep”) to 10 (“pain completely interferes with sleep [unable to sleep due to pain]”).

Secondary outcome measures at randomization and endpoint included the patient-reported Medical Outcomes Study-Sleep Scale (MOS-SS) [17], Brief Pain Inventory-Short Form (BPI-sf) [18], European Quality of Life 5-Dimensions Questionnaire (EQ-5D) [19], Healthcare Utilization Economic Assessment, and Work Productivity and Activity Impairment Questionnaire-Specific Health Problem (WPAI-SHP) [20]. Patients completed the Patient Global Impression of Change (PGIC) using a 7-point scale from “very much worse” to “very much improved”. At endpoint, only the BPI-sf was summarized for the pain severity index score (average of 4 individual pain scores) and pain interference index score (average of 7 individual interference scores).
Statistical analyses

Sample size was calculated using estimates of variance and treatment difference from a previous PTNP trial. A sample of 235 patients per arm would provide 90% power to detect a treatment difference from placebo of 0.6 with respect to change from baseline to week 15 in mean pain scores (MPS), assuming standard deviation (SD) of 2.0 and type I error rate of 0.05. A preplanned, unblinded interim analysis to re-estimate the sample size up to a maximum of 700 patients was performed by an independent data and safety monitoring board when approximately 80% of 470 patients had completed or discontinued from the study. The interim analysis was not intended to stop the study early for any efficacy claim.

The prespecified primary analysis compared change from baseline to week 15 in the MPS. The primary analysis employed mixed-model repeated measures (MMRM), using SAS PROC MIXED with model terms of treatment, trauma type, center, week, treatment-by-week interaction, and baseline MPS as covariates. Week was used as a class variable. A multiple imputation (MI) method was used to impute missing MPS for the primary efficacy analysis, with poor outcomes imputed for withdrawals due to adverse events (AEs) or lack of efficacy.

Secondary analyses compared the change from baseline in weekly MPS and weekly mean sleep interference score with an MMRM model using model terms of treatment, trauma type, center, week, treatment-by-week interaction, and baseline MPS as covariates, without use of an MI algorithm. For sensitivity analyses, analysis of covariance (ANCOVA) with model terms of treatment, center, trauma type, and baseline value analysis was applied to the change from baseline to week 15 in MPS with imputation methods of baseline observation carried forward (BOCF), last observation carried forward (LOCF), and modified baseline observation carried forward (mBOCF), which applied the BOCF rule for patients discontinued due to AEs and the LOCF rule for patients discontinued for any other reason. Other continuous secondary endpoints were also analyzed with this ANCOVA model.

Responder analyses compared the percentage of participants who achieved a 30% and 50% reduction in MPS from baseline to weeks 1–15 using a generalized linear model with a logistic link function. The model included categorical effects of treatment, center, trauma type, week, and treatment-by-week interaction, as well as a continuous baseline MPS. PGIC was analyzed at the endpoint using the Cochran–Mantel–Haenszel (CMH) test stratified for center and trauma type.

A p value < 0.05 was considered statistically significant for all analyses. Adjustments for multiple secondary analyses were not made. The intention-to-treat (ITT) population, defined as all randomized patients who received at least one dose of study drug, was used for all analyses. A per-protocol population was defined as a subset of the ITT population who completed the study and did not have major protocol deviations with the potential to affect the primary efficacy analysis. The per-protocol population was used only in a sensitivity analysis of the primary efficacy endpoint.

Results

Demographic and baseline characteristics

This study was conducted between 31 October 2012 and 4 August 2015. Of 1164 patients screened for inclusion, 622 did not meet the inclusion criteria, most often because of failure to meet eligibility criteria (n = 584; 93.9%), specifically, the NeP criteria (n = 190). A total of 542 patients were randomized, 3 of whom were not treated (pregabalin, n = 274; placebo, n = 265) (Fig. 1).

Approximately half the patients were male; mean age was 53 years (range, 20–85 years). Most patients were white (79.7%). Approximately half the patients had pain from surgery, and the remainder had pain from other traumatic injuries. All randomized patients had a primary diagnosis of peripheral nerve injury, with the nerves most commonly affected (i.e., amounting to > 5% of cases) as follows: peroneal (10.1%), ulnar (6.6%), sural (6.1%), median (6.0%), sciatic (5.9%), radial (5.6%), lateral cutaneous nerve of the thigh (5.3%), and other (6.3%). Demographic characteristics were similar between treatment groups. Mean duration between symptom onset and enrollment in the trial was 8.0 years in both treatment groups. The MPS from the baseline week of the daily NRS diary was 6.41 in the pregabalin group and 6.54 in the placebo group (Table 1).

Prior exposure to gabapentinoids (40 patients in total) was higher in the placebo group (9.4%) than in the pregabalin group (5.5%). Concomitant pain drug treatments (continuing stable treatment at randomization) were taken by 120 (43.8%) patients in the pregabalin group and 130 (49.1%) patients in the placebo group. The most common concomitant pain drugs were ibuprofen, acetaminophen, and tramadol. The proportions of patients taking antidepressants for any reason were similar between the treatment groups, comprising less than one-tenth of patients in either group, with a smaller subset of patients taking tricyclic antidepressants or serotonin–norepinephrine reuptake inhibitors. Rescue drug treatments were taken by 16 patients in the pregabalin group and by 37 patients in the placebo group.

At the end of the dose optimization period, the mean maintenance dose for evaluable pregabalin-treated patients (n = 256) was 473.7 mg/day. The maintenance dose of pregabalin was 600 mg/day for 150 (58.6%) patients, 450 mg/day...
for 54 (21.1%) patients, 300 mg/day for 28 (10.9%) patients, and 150 mg/day for 24 (9.4%) patients.

**Efficacy**

**Primary analysis**

The primary analysis did not demonstrate a significant difference between groups in the mean change of pain scores from baseline to week 15 [pregabalin vs placebo, −0.22; 95% confidence interval (CI) −0.54 to 0.10; \( p = 0.1823 \)]. At week 15, pain scores in both groups had improved compared with baseline [least-squares (LS) mean change from baseline: pregabalin, −2.12; placebo, −1.90]. The results were similar when missing data were imputed using sensitivity analyses LOCF, BOCF, and mBOCF.

**Weekly assessments** At all weeks, the pregabalin-treated group had greater improvement in weekly MPS than the placebo-treated group, although a relative increase in placebo response during the final 2 weeks, 14 and 15, was noted. Differences were statistically significant (\( p < 0.05 \)) from week 2 to 13, except for week 5 (Fig. 2). In addition, the overall mean was also statistically significant (\( p < 0.05 \), with a treatment difference of −0.31.

The 30 and 50% responder status was defined for each patient based on the percentage change in the MPS from baseline to each week, 1–15. Overall, there were more pregabalin 30% and 50% pain responders when compared with the placebo group. The incidence of responders generally
increased over the first 3 weeks of treatment for both 30 and 50% responders. Responder data for weeks 1–3, 14, and 15 are summarized for brevity (responder percentage [week, number responders/number observed at week, \( p \) value pregabalin vs placebo]). Pregabalin 30% responders were 11.9% (week 1, 31/260; \( p = 0.0028 \)), 27.2% (week 2, 69/254; \( p = 0.0360 \)), 38.9% (week 3, 98/252; \( p = 0.0235 \)), 57.4% (week 14, 128/223; \( p = 0.4245 \)), and 57.7% (week 15, 113/196; \( p = 0.8464 \)). Placebo 30% responders were 5.0% (week 1, 13/258), 20.1% (week 2, 49/244), 30.2% (week 3, 74/245), 54.3% (week 14, 113/208), and 58.3% (week 15, 109/187). Pregabalin 50% responders were 4.6% (week 1, 12/260; \( p = 0.1633 \)), 11.4% (week 2, 29/254; \( p = 0.0652 \)), 22.6% (week 3, 57/252; \( p = 0.0039 \)), 37.7% (week 14, 84/223; \( p = 0.0314 \)), and 39.8% (week 15, 78/196; \( p = 0.1889 \)). Placebo 50% responders were 2.3% (week 1, 6/258), 7.0% (week 2, 17/244), 13.5% (week 3, 33/245), 29.8% (week 14, 62/208), and 34.2% (week 15, 64/187). Differences were statistically significant for pregabalin compared with placebo (\( p < 0.05 \), generalized linear model) for 30% responders at weeks 1–3, and for 50% responders at weeks 3–5, 7, 9, and 11–14. In the BPI-sf mean pain severity index, a statistically significant change from baseline to week 15 favored pregabalin over placebo (\( p = 0.0050 \); ANCOVA). The LS mean (standard error) change was −2.40 (0.13) in the pregabalin group and −1.95 (0.13) in the placebo group. In the mean pain interference index, a statistically significant change from baseline also favored pregabalin over placebo (\( p = 0.0168 \)) (Table 2).

At week 15, the PGIC ratings of improvement were statistically significantly higher in the pregabalin group than the placebo group. More patients in the pregabalin group (\( n = 157; 60.9\% \)) than the placebo group (\( n = 120; 48.2\% \)) reported that they were very much or much improved on the PGIC (\( p = 0.0029 \), CMH). Two additional predefined approaches for PGIC analysis, involving different groupings of the categories of improvement (“very much” or “much improved” vs all other groups; and categories comprising “any improvement”, “no change”, “any worsening”), also demonstrated statistically significant results favoring pregabalin.

Sleep The weekly mean sleep interference score at week 15 showed significantly greater improvement in the pregabalin group than in the placebo group [difference in

| Table 1 Baseline characteristics (randomized population) |
|--------------------------------------------------------|
| Demographic characteristics | Pregabalin (\( n = 275 \)) | Placebo (\( n = 267 \)) | Total (\( N = 542 \)) |
| Age (years) | 52.8 (12.9) | 53.5 (12.6) | 53.1 (12.8) |
| Age, years, no. (%) of patients | | | |
| 18–44 | 60 (21.8) | 61 (22.8) | 121 (22.3) |
| 45–64 | 163 (59.3) | 156 (58.4) | 319 (58.9) |
| ≥ 65 | 52 (18.9) | 50 (18.7) | 102 (18.8) |
| Race, no. (%) of patients | \( N = 274 \) | \( N = 265 \) | | |
| White | 217 (78.9) | 215 (80.5) | 432 (79.7) |
| Black | 47 (17.1) | 46 (17.2) | 93 (17.2) |
| Asian | 7 (2.5) | 3 (1.1) | 10 (1.8) |
| Other | 4 (1.5) | 3 (1.1) | 7 (1.3) |
| Weight (kg) | Mean (SD) | 85.9 (20.0) | 86.2 (19.0) | 86.1 (19.5) |
| Height (cm) | Mean (SD) | 170.1 (10.1) | 168.6 (9.4) | 169.4 (9.8) |
| Trauma type, no. (%) of patients | Surgical | 131 (47.6) | 138 (51.7) | 269 (49.6) |
| Baseline mean pain (daily NRS) | | | |
| Mean (SD) | 6.41 (1.3) | 6.5 (1.3) | | |
| Questionnaire: PainDETECT [13] | Mean | 23.1 (5.52) | 22.6 (5.52) | | |
| Range | 2–36 | 3–38 | | |

\( NRS \) numeric rating scale, \( SD \) standard deviation
adjusted mean change (pregabalin vs placebo) − 0.43; 95% CI, − 0.71 to − 0.14; *p = 0.0031*. Adjusted mean change in sleep interference from baseline to week 15 in the pregabalin and placebo groups was − 2.29 (0.11) and − 1.86 (− 0.11), respectively. Statistically significant differences favoring
pregabalin were observed at all weeks (Supplementary Fig. S2).

Safety

Pregabalin was well tolerated in this study, and the safety profile was consistent with or demonstrated fewer AEs when compared to the known profile of pregabalin in other NeP conditions [9]. Adverse events, most frequently dizziness and somnolence, occurred more frequently in the pregabalin group than the placebo group (Table 3). All pregabalin dose levels used in the study were combined for reporting AEs consistent with a flexible dose design in which doses were titrated based on patients’ responses and treatment tolerability. Overall, 138 (50.4%) patients in the pregabalin group and 106 (40.0%) in the placebo group experienced at least one AE. Two patients in the pregabalin group and seven in the placebo group experienced serious adverse events (SAEs), none of which was considered related to treatment. One death occurred in the placebo group. There were no clinically significant findings with respect to changes in laboratory values, vital signs, physical and neurologic examinations, or electrocardiogram.

There were 53 (19.3%) and 16 (6.0%) patients in the pregabalin and placebo groups, respectively, who had dose reductions or temporary discontinuation of treatment due to AEs. Treatment-emergent (all-causality) AEs led to permanent discontinuation from the study for 13 (4.7%) patients in the pregabalin group and 15 (5.7%) patients in the placebo group. Dizziness led to permanent discontinuation of treatment for two patients in the pregabalin group and one in the placebo group. Discontinuation for somnolence occurred for two patients in the pregabalin group.

Discussion

This randomized, double-blind, placebo-controlled, parallel-group study of patients with PTNP did not demonstrate a statistically significant treatment effect with pregabalin on mean pain intensity scores at the prespecified primary efficacy endpoint of the study, week 15. However, the primary efficacy parameter (change in mean pain intensity scores) was statistically significant at most other time points, and several secondary outcomes related to analgesic efficacy and quality of life improved with pregabalin compared with placebo treatment. Pregabalin was superior to placebo in reducing pain severity and interference with daily function as measured by the BPI-sf. A significantly greater number of patients in the pregabalin group described their overall status as very much improved or much improved when comparing their assessment on the PGIC during the last week of treatment (i.e., week 15) to that at study outset. Finally, there was a consistent and significant reduction in pain interference with sleep over the course of the study.

This was the first large phase 3, randomized, controlled trial designed to evaluate the analgesic efficacy of pregabalin in PTNP for registration purposes in the United States. The conduct of this trial indicates that a large, multinational, phase 3 trial of pregabalin in PTNP is feasible. The significant effect of pregabalin on a number of secondary endpoints, including second measure of pain intensity (BPI-sf) raises the possibility that the absence of a statistically significant difference on the primary endpoint may be partly attributed to additional factors individually or in aggregate including (1) the study design, (2) dose titration, (3) adequacy of dose achieved, (4) duration of the fixed dose period, (5) the number and wide geographical location of study sites needed to conduct the study, (6) the level of refractoriness of the study population, and (7) differential response to study treatments over time in study, particularly the placebo response. A review of randomized, parallel-group, placebo-controlled trials in neuropathic pain found that the magnitude of the placebo response accrues slowly and continually, whereas pain reduction in response to active treatments manifests more rapidly, before leveling off [21]. Such a pattern of placebo response occurred over the course of the present study and possibly explains why pregabalin demonstrated a significant treatment effect on mean pain intensity in earlier weeks, but not in the primary analysis at week 15. Moreover, the efficacy of pregabalin observed in the earlier weeks of this trial is consistent with the results from a prior, shorter

Table 3

| Adverse events (treatment related) by preferred term | No. (%) of patients with adverse events (N=274) | Placebo (N=265) |
|---------------------------------------------------|-----------------------------------------------|-----------------|
| Dizziness                                           | 40 (14.6)                                    | 11 (4.2)        |
| Somnolence                                          | 27 (9.9)                                     | 9 (3.4)         |
| Fatigue                                             | 14 (5.1)                                     | 10 (3.8)        |
| Nausea                                              | 14 (5.1)                                     | 8 (3.0)         |
| Headache                                            | 12 (4.4)                                     | 8 (3.0)         |
| Vertigo                                             | 12 (4.4)                                     | 1 (0.4)         |
| Back pain                                           | 8 (2.9)                                      | 5 (1.9)         |
| Disturbance in attention                            | 8 (2.9)                                      | 0               |
| Memory impairment                                   | 7 (2.6)                                      | 0               |
| Nasopharyngitis                                     | 7 (2.6)                                      | 8 (3.0)         |
| Constipation                                        | 6 (2.2)                                      | 4 (1.5)         |
| Sedation                                            | 6 (2.2)                                      | 0               |
| Pain in extremity                                   | 2 (0.7)                                      | 6 (2.3)         |
| Insomnia                                            | 2 (0.7)                                      | 6 (2.3)         |

*Includes data up to 999 days after last dose of study drug. Medical Dictionary for Regulatory Activities (MedDRA) v18.1 coding dictionary applied*
trial with an 8-week double-blind period [10]. Although subjects in this study by Van Seventer et al. had an equally diverse set of etiologic mechanisms associated with chronic neuropathic pain [10], it should also be noted that the duration of the syndromes was nearly 50% shorter (4.4 vs 8 years in this study) and, therefore, the studies may not be strictly comparable.

In future studies designed for a post-traumatic neuropathic pain study population, a few considerations are worth noting. For example, a retrospective analysis demonstrated the effect size was larger in the post-surgical vs non-surgical subgroup (−0.43 vs 0.05, respectively), and nominal unadjusted $P$ value was 0.04 compared with placebo (data on file). However, this is caveated in that the surgical and non-surgical subgroup analyses were not preplanned prior to unblinding, were conducted post hoc and not adjusted for multiple comparisons. It may be, for example, that post-surgical traumatic pain contrasted with non-surgical pain is diagnosed with greater definitiveness and/or that such subjects can rate their pain and change in pain better than non-surgical subjects. Thus, it would be premature to suggest meaningful significance to this finding; however, this finding could be considered or explored in future research. Fixed-dose studies in contrast to flexible dosing in registration studies might help to address any potential concerns for the adequacy of dose studied. When feasible, using countries with a well-established research infrastructure for neuropathic pain research may be advisable. Performing the final assessment 1 week prior to the end of the study while not making this explicitly apparent to the subjects may potentially avoid the presumed anticipatory narrowing of the treatment difference observed at the end of the current study.

The tolerability of pregabalin observed in this study was consistent with the known profile of pregabalin in the management of NeP [22, 23]. Adverse events were preponderantly dizziness, somnolence, and fatigue, which generally resolved as treatment continued. Treatment-emergent AEs of all-causality resulted in permanent withdrawal from the study for 13 patients in the pregabalin group and 15 in the placebo group. The two SAEs that occurred among pregabalin-treated patients were not ascribed to the study drug.

A few limitations warrant consideration. The methodology included the most recent recommendations intended to enhance assay sensitivity; some of these have face validity but have yet to be empirically tested [12, 24]. A related limitation was the lack of capacity to determine the effect of the recommended methods on assay sensitivity. Many variables related to study design influence assay sensitivity making it difficult to draw conclusions from a single study. The acute mechanisms of nerve injury within the enrolled population were etiologically diverse, and it may be difficult to discern in some cases, whether a peripheral nerve injury and corresponding deficits characterized prior to surgery were not worsened by tissue manipulation, electrocautery, and/or incisions of the required surgery. Eighteen percent of participants did not complete the study, and the corresponding loss of data may have affected the primary result. Treatment-emergent side effects of pregabalin had the potential for partial unblinding, which may have led to bias. Finally, performing multiple analyses and basing conclusions of statistical significance on $p$ values <0.05 may lead to false-positive inferences; however, considering that the preponderance of clinically relevant secondary analyses yielded $p$ values <0.05, the totality of the data appears to support a clinical benefit for pregabalin in PTNP.

The primary analysis of this trial was not statistically significant; however, secondary analyses of multiple outcome measures were. Statistically significant changes were relatively modest, and there is no consensus on the magnitude of a group difference for the 0–10 NRS that is considered clinically meaningful. The efficacy of pregabalin in PTNP merits further study in light of promising findings in clinically relevant outcomes, the low rate of SAEs, the positive results of a previous trial, established efficacy in other chronic NeP syndromes, and the lack of evidence-based treatments for PTNP.

**Key points**

**Question** Is pregabalin efficacious and tolerable for the treatment of chronic, post-traumatic neuropathic pain?

**Findings** In a double-blind, randomized international study of 542 evaluable patients (pregabalin $n = 274$) of whom approximately half had post-surgical neuropathic pain, the primary efficacy analysis did not demonstrate a statistically significant difference between active treatment and placebo in change from baseline to week 15 ($p = 0.1823$). However, comparisons for key secondary outcome measures yielded $p$ values <0.05 favoring pregabalin. Safety and tolerability were consistent with the known profile of pregabalin.

**Meaning** Additional studies are needed to characterize the efficacy and tolerability of pregabalin for chronic, post-traumatic neuropathic pain.

**Acknowledgements** This study was sponsored by Pfizer Inc. None of the authors were paid to write this article, and all authors had full access to all data in the study. Medical writing support was provided by Karen Burrows and Rosemary Perkins of Engage Scientific Solutions and was funded by Pfizer. All authors gave final consent for submission of this article.

**Compliance with ethical standards**

**Ethical standards** The protocol complied with the Declaration of Helsinki (1964), and was reviewed and approved by the institutional review board at each participating center.
Informed consent  All participants provided written, informed consent.

Conflicts of interest  JM has participated in advisory boards (Pfizer, Editas Medicine, Flexion Therapeutics, Teva, Quark, Pacira, Inspiration on Delivery Sciences, Quartet, Pacira Egalet, Biogen, Nektar, Endo, Immune Pharma, Chromocell, Collegium, Purdue, Novartis, Sanofi, Convergence, Aptinyx, Daiichi Sankyo, Allergan, Plasmaphysurgical, and Grunenthal), received research funding (Depomed, Pfizer), and served on Data Safety Monitoring Boards (Allergan, Novartis). M Resnick, R Yang, J Scavone, E Whalen, G Gregorian, B Parsons, and L Knapp are employees of Pfizer Inc. S Greenberg was an employee of Pfizer at the time of the study and development of the manuscript. N Katz is employed by Analgesic Solutions, which provided a central eligibility verification service to identify patients with PTNP.

Data sharing statement  Upon request, and subject to certain criteria, conditions and exceptions (see https://www.pfizer.com/science//clinical-trials/trial-data-and-results for more information), Pfizer will provide access to individual de-identified participant data from Pfizer-sponsored global interventional clinical studies conducted for medicines, vaccines and medical devices (1) for indications that have been approved in the US and/or EU or (2) in programs that have been terminated (i.e., development for all indications has been discontinued). Pfizer will also consider requests for the protocol, data dictionary, and statistical analysis plan. Data may be requested from Pfizer trials 24 months after study completion. The de-identified participant data will be made available to researchers whose proposals meet the research criteria and other conditions, and for which an exception does not apply, via a secure portal. To gain access, data requestors must enter into a data access agreement with Pfizer.

Open Access  This article is distributed under the terms of the Creative Commons Attribution 4.0 International License (http://creativecommons.org/licenses/by/4.0/), which permits unrestricted use, distribution, and reproduction in any medium, provided you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons license, and indicate if changes were made.

References

1. Crombie IK, Davies HT, Macrae WA (1998) Cut and thrust: ante-cedent surgery and trauma among patients attending a chronic pain clinic. Pain 76(1–2):167–171
2. Kehelet H, Jensen TS, Woolf CJ (2006) Persistent postoperative pain: risk factors and prevention. Lancet 367(9522):1618–1625. https://doi.org/10.1016/S0140-6736(06)68700-X
3. Treede RD, Jensen TS, Campbell JN, Cruccu G, Dostrovsky JO, Griffin JW, Hansson P, Hughes R, Nurmikko T, Serra J (2008) Neuropathic pain: redefinition and a grading system for clinical and research purposes. Neurology 70(18):1630–1635. https://doi.org/10.1212/0w1.01200.0000287263.29778.59
4. Haanpaa M, Attal N, Backonja M, Baron R, Bennett M, Bouhas-sira D, Cruccu G, Hansson P, Haythornwaite JA, Iannetti GD, Jensen TS, Kauppi A, Nurmikko TJ, Rice AS, Rowbotham M, Serra J, Sommer C, Smith BH, Treede RD (2011) NeuPSIG guidelines on neuropathic pain assessment. Pain 152(1):14–27. https://doi.org/10.1016/j.pain.2010.07.031
5. von Hehn CA, Baron R, Woolf CJ (2012) Deconstructing the neuropathic pain phenotype to reveal neural mechanisms. Neuron 73(4):638–652. https://doi.org/10.1016/j.neuron.2012.02.008
6. Gilron I, Kehelet H (2014) Prevention of chronic pain after surgery: new insights for future research and patient care. Can J Anaesth 61(2):101–111. https://doi.org/10.1007/s12630-013-0067-8
7. Macrae WA (2008) Chronic post-surgical pain: 10 years on. Br J Anaesth 101(1):77–86. https://doi.org/10.1093/bja/aen099
8. Wong K, Phelan R, Kalso E, Galvin I, Goldstein D, Raja S, Gilron I (2014) Antidepressant drugs for prevention of acute and chronic post-surgical pain: early evidence and recommended future directions. Anesthesiology 121(3):591–608. https://doi.org/10.1097/ALN.0000000000000307
9. Lyrica (pregabalin) [Prescribing Information]. Pfizer Inc. NY, NY (2016) http://labeling.pfizer.com/ShowLabeling.aspx?id=561. Accessed 25 Jan 2018
10. van Severen R, Bach FW, Toth CC, Serpell M, Temple J, Murphy TK, Nimour M (2010) Pregabalin in the treatment of post-traumatic peripheral neuropathic pain: a randomized double-blind trial. Eur J Neurol 17(8):1082–1089. https://doi.org/10.1111/j.1468-1331.2010.02979.x
11. US Department of Health and Human Services FaDA, Center for Drug Evaluation and Research (CDER) (2014) Guidance for industry analgesic indications: developing drug and biological products. http://www.fda.gov/downloads/Drugs/GuidanceComplianceRegulatoryInformation/Guidances/UCM384691.pdf. Accessed 10 Nov 2017
12. Dworkin RH, Turk DC, Peirce-Sandner S, Burke LB, Farrar JT, Gilron I, Jensen MP, Katz NP, Raja SN, Rappaport BA, Rowbotham MC, Backonja MM, Baron R, Bellamy N, Bhagwagar Z, Costello A, Cowan P, Fang WC, Hertz S, Jay GW, Junor R, Kerns RD, Ker-win R, Kopecky EA, Lissin D, Malamut R, Markman JD, McDermott MP, Munera C, Porter L, Raenscholk C, Rice AS, Sampaio C, Skljarevski V, Sommerville K, Stacey BR, Steigerwald I, Tobias J, Treintcosti AM, Wasan AD, Wells GA, Williams J, Witter J, Ziegler D (2012) Considerations for improving assay sensitivity in chronic pain clinical trials: IMMPACT recommendations. Pain 153(6):1148–1158. https://doi.org/10.1016/j.pain.2012.03.003
13. Freynhagen R, Baron R, Gockel U, Tolle TR (2006) painDETECT: a new screening questionnaire to identify neuropathic components in patients with back pain. Curr Med Res Opin 22(10):1911–1920. https://doi.org/10.1185/030079906X132488
14. Burstein R, Collins B, Jakubowski M (2004) Defeating migraine pain with triptans: a race against the development of cutaneous allo-dynia. Ann Neurol 55(1):19–26. https://doi.org/10.1002/ana.10786
15. Kroenke K, Spitzer RL, Williams JB (2002) The PHQ-15: validity of a new measure for evaluating the severity of somatic symptoms. Psychosom Med 64(2):258–266
16. Posner K, Brent D, Lucas C, Gould M, Stanley B, Brown G, Fisher P, Zelazny J, Burke A, Oquendo M, Mann J (2009) Columbia-suicide severity rating scale (C-SSRS) The Research Foundation for Mental Hygiene, New York
17. Hays RD, Stewart A (1992) Measuring functioning and well-being: the medical outcomes study approach. Duke University Press, Durham
18. Cleeland CS (1985) Measurement and prevalence of pain in cancer. Semin Oncol Nurs 1:87–92
19. EuroQol Group (1990) EuroQo—a new facility for the measurement of health-related quality of life. Health Policy 16:199–208
20. Reilly MC, Zbrozek AS, Dukes EM (1993) The validity and reproducibility of a work productivity and activity impairment instrument. Pharmacoeconomics 4(5):353–365
21. Quessy SN, Rowbotham MC (2008) Placebo response in neuro-pathic pain trials. Pain 138(3):479–483. https://doi.org/10.1016/j. pain.2008.06.024
22. Rosenstock J, Tuchman M, LaMoreaux L, Sharma U (2004) Pregabalin for the treatment of painful diabetic peripheral neuropathy: a double-blind, placebo-controlled trial. Pain 110(3):628–638. https://doi.org/10.1016/j.pain.2004.05.001
23. Sabatowski R, Galvez R, Cherry DA, Jacquot F, Vincent E, Mai-sonobe P, Versavel M, Study G (2004) Pregabalin reduces pain and improves sleep and mood disturbances in patients with post-herpetic
neuralgia: results of a randomised, placebo-controlled clinical trial. Pain 109(1–2):26–35. https://doi.org/10.1016/j.pain.2004.01.001

24. Demant DT, Lund K, Vollert J, Maier C, Segerdahl M, Finnerup NB, Jensen TS, Sindrup SH (2014) The effect of oxcarbazepine in peripheral neuropathic pain depends on pain phenotype: a randomised, double-blind, placebo-controlled phenotype-stratified study. Pain 155(11):2263–2273. https://doi.org/10.1016/j.pain.2014.08.014