Eigenvalues and Singular Values of Products of Rectangular Gaussian Random Matrices

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We derive exact analytic expressions for the distributions of eigenvalues and singular values for the product of an arbitrary number of independent rectangular Gaussian random matrices in the limit of large matrix dimensions. We show that they both have power–law behavior at zero and determine the corresponding powers. We also propose a heuristic form of finite size corrections to these expressions which very well approximates the distributions for matrices of finite dimensions.

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I. INTRODUCTION

Spectral analysis of the products of random matrices is a powerful tool in several domains of statistical physics, allowing, for example, to study Lyapunov exponents for disordered and chaotic dynamical systems [1]. It is also useful in a class of problems related to multiplicative matrix–valued noncommutative diffusion processes [2]. Several applications go beyond physics, as for instance, those related to the stability analysis of ecological systems [3] or to telecommunication applications based on the scattering of electromagnetic waves on random obstacles [4, 5]. In many of these cases, some exact analytic results were obtained for relatively small matrices. Interestingly, quite often analytic calculations are possible under another limit — the limit of matrix dimensions tending to infinity. Examples include products of pseudounitary matrices, representing transfer matrices in mesoscopic wires [6], large \( N \) Wilson loops in Yang–Mills theory [7, 8] or multiplicative diffusion of infinitely large complex and/or Hermitian matrices \[9, 10, 11\]. In most of these cases, the reason why the exact spectral distribution is within the reach of analytic methods is due to a link to free random variable calculus \[12, 13\], which is a very powerful technique. This is also why the spectra of products of large random matrices represent a challenge for mathematicians \[14, 15\]. In this paper, we generalize the analysis of the product of large, square, random Gaussian matrices, performed in \[16\], to the product of rectangular matrices. In particular, we study the product

\[
P = A_1 A_2 \ldots A_L
\]  

(1)

of \( L \geq 1 \) independent, rectangular, large, random Gaussian matrices \( A_l, l = 1, 2, \ldots, L \), of dimensions \( N_l \times N_{l+1} \). We are interested in the eigenvalue and singular value density of \( P \) in the limit \( N_{L+1} \to \infty \) and

\[
R_l = \frac{N_l}{N_{L+1}} = \text{finite}, \quad \text{for} \quad l = 1, 2, \ldots, L + 1.
\]  

(2)
In other words, all matrix dimensions grow to infinity at fixed rates and, obviously, $R_{L+1} = 1$. The product $P$ is a matrix of dimensions $N_1 \times N_{L+1}$ and has eigenvalues only if it is a square matrix: $N_1 = N_{L+1}$.

We assume the matrices $A_i$ in the product (1) to be complex Gaussian matrices drawn randomly from the ensemble defined by the probability measure

$$
d\mu(A_i) \propto e^{-dN_{N_{i+1}} \sigma_f^2 \text{Tr}(A_i^* A_i)} \, dA_i,
$$

where $dA = \prod_{a,b} d(\text{Re}[A_{ab}]) d(\text{Im}[A_{ab}])$ is a flat measure. A normalization constant, fixed by the condition $\int d\mu(A) = 1$, is omitted. This is the simplest generalization of the Girko–Ginibre ensemble [18–20] to rectangular matrices. The $\sigma_i$ parameters set the scale for the Gaussian fluctuations in $A_i$’s. The entries of each matrix $A_i$ can be viewed as independent centered Gaussian random variables, the variance of the real and imaginary parts being proportional to $\sigma_i^2$ and inversely proportional to the square root of the number $N_i N_{i+1}$ of elements in the matrix.

The eigenvalue density of the product (1) of square Gaussian matrices was calculated in [16] while the singular value distribution was determined in [14, 15, 17]. The eigenvalue density was obtained using Free Random Variables calculus [13, 24, 25]. Both techniques work in the infinite matrix size limit. After explaining notation (Section II) and listing the main results of the paper (Section III), we shall follow those same methods to derive the corresponding results for the product of rectangular matrices. In Section IV, we present a diagrammatic derivation of the moment generating function for the product $P$. In Section V, using the tools of Free Random Variables calculus, we obtain the moment generating function for $Q = P^\dagger P$, recovering results given in [17]. Section VI concludes the paper with a discussion on particular applications of our results and possible generalizations.

### II. Generalities

Let us spend a few words on the notations to be used in this paper. The eigenvalue density $\rho_X(\lambda)$ of a Hermitian matrix $X$ is a real function of real argument, while in the case of a non–Hermitian matrix it is a real function of complex argument. In the latter case we shall write $\rho_X(\lambda, \bar{\lambda})$ and treat $\lambda$ and its conjugate $\bar{\lambda}$ as independent variables.

In the Hermitian case, the eigenvalue density can be computed from a Green’s function $G_X(z)$ [26, 27] which contains the same information as the density itself:

$$
\rho_X(\lambda) = -\frac{1}{\pi} \lim_{\epsilon \to 0^+} \text{Im} G_X(\lambda + i\epsilon).
$$

For a non–Hermitian matrix, the corresponding Green’s function $G_X(z, \bar{z})$ is non–holomorphic and therefore we shall write it explicitly as a function of $z$ and $\bar{z}$. In this case the eigenvalue distribution is reconstructed from the Green’s function as [28, 30]

$$
\rho_X(z, \bar{z}) = \frac{1}{\pi} \frac{\partial}{\partial \bar{z}} G_X(z, \bar{z}).
$$

Actually, this equation reduces to (4) when the non–holomorphic region shrinks to a cut along the real axis, as it happens for Hermitian matrices. The Green’s function $G_X(z)$ for a Hermitian matrix is written as a function of a single argument since everywhere except on the cut one has $\partial \bar{z} G_X = 0$, and thus it is $\bar{z}$–independent.

In many applications it is often convenient to use the moment generating function, or $M$–transform, which is closely related to the Green’s function: $M_X(z) = zG_X(z) - 1$. For a Hermitian matrix $X$ one has

$$
M_X(z) = \sum_{n \geq 1} \frac{m_n}{z^n} = \sum_{n \geq 1} \frac{1}{z^n} \int \rho_X(\lambda) \lambda^n d\lambda,
$$

where the $m_n$’s are the moments of the eigenvalue density. If the matrix $X$ is of finite dimensions $N \times N$, the moments are given by $m_n = \text{Tr}(X^n)$. The moment generating function encodes the same information as the Green’s function $G_X(z) = z^{-1} M_X(z) + z^{-1}$. Thus, one can calculate the corresponding eigenvalue distribution from $M_X(z)$.

One can also introduce a similar function for non–Hermitian matrices: $M_X(z, \bar{z}) = zG_X(z, \bar{z}) - 1$. In this case, however, it does not play the role of a moment generating function anymore, since now one also has mixed moments $\langle \text{Tr} X^n (X^\dagger)^k \rangle$, which in general depend on the ordering of $X$ and $X^\dagger$ in the product under the trace.
The situation is slightly simplified when the \( M \)-transform is a spherically symmetric function: \( M_X(z, \bar{z}) = M_X(|z|^2) \). In this case equation (5) can be cast into the form

\[
\rho_X(z, \bar{z}) = \frac{1}{\pi} M_X(|z|^2) + f \delta^2(z, \bar{z})
\]  

(7)

where \( M'_X \) is the first derivative of \( M_X \) and \( f = 1 + M_X(0) \) is a constant representing the fraction of zero modes. In this case, the eigenvalue distribution is spherically symmetric as well (see for example [16] for the product of square matrices). As we shall see later, this is also going to be the case for the product (1) of rectangular Gaussian matrices (3).

III. RESULTS

The matrix \( P (1) \) has eigenvalues only if it is square, while it has singular values for any rectangular shape. As a matter of fact, its singular values can be determined as the square roots of the non–zero eigenvalues of the matrix

\[
Q ≡ P^\dagger P
\]  

(8)

or, alternatively, of the matrix \( R = PP^\dagger \). \( Q \) and \( R \) are Hermitian, and they have non–negative spectra which differ only in the zero modes.

The main finding of this paper is that the eigenvalue distribution and the \( M \)-transform of the product (1) are spherically symmetric. We shall show the \( M \)-transform to satisfy the \( L \)-th order polynomial equation

\[
\prod_{l=1}^{L} \left( \frac{M_P(|z|^2)}{R_l} + 1 \right) = \frac{|z|^2}{\sigma^2}.
\]  

(9)

where the scale parameter is \( \sigma = \sigma_1 \sigma_2 \ldots \sigma_L \). When all of the matrices involved are square, this equation reproduces the results in [16].

An analogous equation for \( Q \) reads

\[
\sqrt{R_1} M_Q(z) + 1 \prod_{l=1}^{L} \left( \frac{M_Q(z)}{R_l} + 1 \right) = \frac{z}{\sigma^2}.
\]  

(10)

It is an algebraic equation of order \((L+1)\), and it was first obtained in [17] in the context of wireless telecommunication. Equations (9) and (10) are strikingly similar. They actually differ only by the prefactor in front of the product. Moreover, the free argument in the first equation is \( |z|^2 \), while \( z \) in the second one. This observation represents the second main result of this paper. Since conventions used in telecommunication theory and in physics differ a bit, in Section V we rederive equation (10) for completeness.

When \( P \) is a square matrix, then \( R_l = 1 \) and the square root at the beginning of equation (10) can be omitted. When the product of square matrices is considered, all of the \( R_l \)’s become equal to unity and the two equations take the following form:

\[
(M_P(|z|^2) + 1)^L = \frac{|z|^2}{\sigma^2}, \quad M_Q^{-1}(z) (M_Q(z) + 1)^{L+1} = \frac{z}{\sigma^2}.
\]  

(11)

Equations (9) can be easily rewritten in terms of the corresponding Green’s functions (see the previous section). If one does that and then applies the prescriptions in (5) and (4) respectively, it becomes clear that

\[
\rho_P(\lambda, \bar{\lambda}) \sim |\lambda|^{-2\frac{L-1}{L}}, \quad \text{and} \quad \rho_Q(\lambda) \sim \lambda^{-\frac{L+1}{L}}, \quad \text{as} \quad \lambda \to 0.
\]  

(12)

In the more general case of rectangular matrices, when solving equations (9) and (10) for the Green’s functions, one can then see that only those brackets in which \( R_l = 1 \) contribute to the singularity at zero, while all others approach a constant for \( z \to 0 \). Thus, the eigenvalue density displays the following singularity

\[
\rho_P(\lambda, \bar{\lambda}) \sim |\lambda|^{-2\frac{L}{L}}, \quad \text{as} \quad \lambda \to 0,
\]  

(13)
where \( s \) is the number of those ratios among \( R_1, \ldots, R_L \) which are exactly equal to unity. On the other hand, the eigenvalue density of \( Q \) behaves as

\[
\rho_Q(\lambda) \sim \lambda^{-\frac{d+1}{2}}, \quad \text{as} \quad \lambda \to 0. \tag{14}
\]

The third result we want to mention here is a heuristic form for the finite size corrections to the eigenvalue distribution. For a large but finite order of magnitude \( N \) of the matrices involved, the eigenvalue distribution is still spherically symmetric. So, let \( \rho_N(r) \) denote the radial profile of this distribution, where \( r = |\lambda| \). As we shall show, the evolution of the radial shape with the size \( N \) is very well described by a simple multiplicative correction:

\[
\rho_N(r) \equiv \rho(r) \frac{1}{2} \text{erfc} \left( q(r - \sigma) \sqrt{N} \right). \tag{15}
\]

In the \( N \to \infty \) limit the correction becomes a step function, so that \( \rho_\infty(r) = \rho(r) \) for \( r \leq \sigma \) and \( \rho_\infty(r) = 0 \) otherwise. The shape of the limiting radial distribution \( \rho(r) \) comes from the solution of (9). This type of finite size corrections can be derived analytically for Girko–Ginibre matrices [31–33]. Here we show that it also works very well for the eigenvalues of the product of Gaussian matrices. It is very generic and possibly applies to other random matrix ensembles with spherically symmetric eigenvalue densities.

IV. THE EIGENVALUES OF A PRODUCT OF RECTANGULAR GAUSSIAN RANDOM MATRICES

In this section, we present a derivation of the main result of this article, equation (9), a realization of it in the case \( L = 2 \), and numerical simulations to confirm our findings. To this end, we employ a technique for summing planar diagrams, the Dyson–Schwinger equations (more precisely described in [10, 16]), extended to a non–Hermitian framework.

The evaluation of the Green’s function for a product \( P \) of random ensembles by means of planar diagrammatics is non–linear w.r.t. the constituent matrices. It is possible to linearize the problem by means of the following trick [10, 16]. Consider the following block matrix:

\[
\tilde{P} = \left( \begin{array}{cccc}
0 & A_1 & 0 & \ldots & 0 \\
0 & 0 & A_2 & \ldots & 0 \\
\vdots & \vdots & \vdots & \ddots & \vdots \\
0 & 0 & 0 & \ldots & A_{L-1} \\
A_L & 0 & 0 & \ldots & 0
\end{array} \right). \tag{16}
\]

It is a matrix of dimensions \( N_{\text{tot}} \times N_{\text{tot}} \), where \( N_{\text{tot}} = N_1 + N_2 + \ldots + N_L \). It is important to notice that the \( L \)-th power of \( \tilde{P} \) is a block diagonal matrix [10, 16]

\[
\tilde{P}^L = \left( \begin{array}{cccc}
B_1 & 0 & 0 & \ldots & 0 \\
0 & B_2 & 0 & \ldots & 0 \\
0 & 0 & B_3 & \ldots & 0 \\
\vdots & \vdots & \vdots & \ddots & \vdots \\
0 & 0 & 0 & \ldots & B_L
\end{array} \right). \tag{17}
\]

with square blocks \( B_1 = A_1A_2 \ldots A_{L-1}A_L, \) \( B_2 = A_2 \ldots A_LA_1, \ldots, \) being cyclicly permuted products of \( A_1, A_2, \ldots, A_L \). All these blocks have identical non–zero eigenvalues. They differ only in the number of eigenvalues, which may vary from block to block. The first diagonal block \( B_1 \) is equal to the product \( P \) [1]. Taking into account that this block has dimensions \( N_1 \times N_1 \), while the total matrix \( \tilde{P}^L \) has dimensions \( N_{\text{tot}} \times N_{\text{tot}} \), one can easily deduce the following relation between the \( M \)-transforms of \( P \) and \( \tilde{P} \):

\[
M_{\tilde{P}}(w, \overline{w}) = \frac{L}{N_{\text{tot}}} M_{P}(w^L, \overline{w^L}). \tag{18}
\]

The importance of this relation relies in the fact that one can use it to calculate \( M_P(z, \overline{z}) \) from \( M_{\tilde{P}}(w, \overline{w}) \). The latter can be calculated using Dyson–Schwinger equations, since the matrix \( \tilde{P} \) is linear w.r.t. the constituent matrices \( A_i \).
The first step in writing the Dyson–Schwinger equations is to know the propagators of the random matrix in question, namely $\tilde{P}$, or more precisely its “duplicated” version:

$$\tilde{P}^D = \begin{pmatrix} \tilde{P} & 0 \\ 0 & \tilde{P}^\dagger \end{pmatrix}.$$  \hspace{1cm} (19)

We shall think of it as a four–block matrix, each block being an $L \times L$ block matrix. We shall denote the $L \times L$ block indices in these four blocks by $lm$ (upper left corner), $\bar{lm}$ (upper right), $\bar{m}n$ (lower left), $\bar{m}\bar{n}$ (lower right), each one covering the range $1, 2, \ldots, L$; for example $(\tilde{P}^D)_{\bar{l}l} = A^1_l$. All the other matrices involved shall inherit this same structure. For instance,

$$G^D = \begin{pmatrix} G^{ww} & G^{w\bar{w}} \\ G^{\bar{w}w} & G^{\bar{w}\bar{w}} \end{pmatrix} = \begin{pmatrix} [G^D]_{11} & [G^D]_{12} & \cdots & [G^D]_{1L} \\ [G^D]_{21} & [G^D]_{22} & \cdots & [G^D]_{2L} \\ \vdots & \vdots & \ddots & \vdots \\ [G^D]_{L1} & [G^D]_{L2} & \cdots & [G^D]_{LL} \end{pmatrix},$$  \hspace{1cm} (20)

and similarly for $W^D$ and $\Sigma^D$ (to be defined in a moment). For the sake of simplicity, we shall disregard some subscripts and symbols of dependence on $w$, $\bar{w}$.

We are interested in computing the Green’s function of $\tilde{P}$, i.e.,

$$G_{\tilde{P}}(w, \bar{w}) = \frac{1}{N_{\text{tot.}}} \text{Tr}_{\bar{w}} G^{ww} = \frac{1}{N_{\text{tot.}}} \sum_{l=1}^L \text{Tr}_{\bar{w}} [G^D]_{ll} L = \frac{1}{N_{\text{tot.}}} \sum_{l=1}^L N_l \mathcal{G}_{ll},$$  \hspace{1cm} (21)

where it is useful to define the normalized traces

$$\mathcal{G}_{ll} = \frac{1}{N_l} \text{Tr}_{\bar{w}} [G^D]_{ll}, \quad \mathcal{G}_{\bar{w}l} = \frac{1}{N_l} \text{Tr}_{\bar{w}} [G^D]_{l\bar{w}}, \quad \mathcal{G}_{\bar{w}\bar{w}} = \frac{1}{N_l} \text{Tr}_{\bar{w}} [G^D]_{\bar{w}l}, \quad \mathcal{G}_{\bar{w}\bar{w}} = \frac{1}{N_l} \text{Tr}_{\bar{w}} [G^D]_{\bar{w}\bar{w}}.$$  \hspace{1cm} (22)

Hence, we should evaluate the $\mathcal{G}_{ll}$’s.

The only non–zero propagators of $\tilde{P}^D$ are readily determined from the probability measures in (3):

$$\langle [\tilde{P}^D]_{ij} \{ [\tilde{P}^D]_{jk} \rangle = \frac{\sigma_1^2}{\sqrt{N_1 N_2}} 1_{N_1} \otimes 1_{N_2},$$

$$\langle [\tilde{P}^D]_{ij} \{ [\tilde{P}^D]_{jk} \rangle = \frac{\sigma_2^2}{\sqrt{N_2 N_3}} 1_{N_2} \otimes 1_{N_3},$$

$$\vdots$$

$$\langle [\tilde{P}^D]_{ij} \{ [\tilde{P}^D]_{jk} \rangle = \frac{\sigma_L^2}{\sqrt{N_L N_1}} 1_{N_L} \otimes 1_{N_1}. \hspace{1cm} (23)$$

Thus, we are now in position to write down the two Dyson–Schwinger equations for $\tilde{P}^D$. The first one, being the definition of the self–energy matrix $\Sigma^D$, is independent of the propagators [10]:

$$G^D = (W^D - \Sigma^D)^{-1},$$  \hspace{1cm} (24)

where $W^D$ is defined as $w1_{N_{\text{tot.}}}$ in its left upper block (where $w \in \mathbb{C}$), $\bar{w}1_{N_{\text{tot.}}}$ in the right lower block, and zero otherwise. The second one is presented in [10, 16], and the structure of the propagators [23] implies that the only non–zero blocks of the self–energy matrix read

$$[\Sigma^D]_{ll} = \frac{\sigma_1^2}{\sqrt{N_l N_{l+1}}} \text{Tr} [G^D]_{l+1,l+1} 1_{N_l} = \sigma_1^2 \sqrt{\frac{N_{l+1}}{N_l}} G_{l+1,l+1} 1_{N_l},$$  \hspace{1cm} (25)
for all \( l = 1, 2, \ldots, L \), with the cyclic convention \( 0 = L \), where the normalized traces \((22)\) have been used.

Results \((23), (26)\) mean that the four blocks of the matrix \((W^D - \Sigma^D)\) are diagonal. Such a matrix can be straightforwardly inverted: its four blocks remain diagonal, and read

\[
(W^D - \Sigma^D)^{-1} = \begin{pmatrix}
\begin{pmatrix}
\bar{w} \gamma_1 \mathbf{1}_{N_1} & 0 & \ldots & 0 \\
0 & \bar{w} \gamma_2 \mathbf{1}_{N_2} & \ldots & 0 \\
\vdots & \vdots & \ddots & \vdots \\
0 & 0 & \ldots & \bar{w} \gamma_L \mathbf{1}_{N_L}
\end{pmatrix}
&
\begin{pmatrix}
\alpha_1 \gamma_1 \mathbf{1}_{N_1} & 0 & \ldots & 0 \\
0 & \alpha_2 \gamma_2 \mathbf{1}_{N_2} & \ldots & 0 \\
\vdots & \vdots & \ddots & \vdots \\
0 & 0 & \ldots & \alpha_L \gamma_L \mathbf{1}_{N_L}
\end{pmatrix}

\end{pmatrix}.
\]

where, for all \( l = 1, 2, \ldots, L \),

\[
\frac{1}{\gamma_l} \equiv |w|^2 - \alpha_l \beta_l = |w|^2 - (\sigma_{l-1} \sigma_l)^2 \frac{\sqrt{N_l-1} N_{l+1}}{N_l} G_{l+1,l+1} G_{l,l-1}^{-1}.
\]

Substituting \((27)\) into \((24)\), we find that the only non–zero blocks of the duplicated Green’s function \((20)\) are, for all \( l = 1, 2, \ldots, L \),

\[
\begin{align*}
[G^D]_{ll} &= \bar{w} \gamma_l \mathbf{1}_{N_l},
\end{align*}
\]

Taking the normalized traces of both sides of every equality in \((29)\) leads to the final set of equations,

\[
\mathcal{G}_{ll} = \bar{w} \gamma_l, \quad \mathcal{G}_{ll} = \alpha_l \gamma_l, \quad \mathcal{G}_{ll} = \beta_l \gamma_l, \quad \mathcal{G}_{ll} = w \gamma_l.
\]

The structure of equations \((30)\) is the following: the fourth one is the conjugate of the first, and it is then redundant. The second and third ones read

\[
\mathcal{G}_{ll} = \sigma_l^2 \sqrt{\frac{N_{l+1}}{N_l}} G_{l+1,l+1} G_{l,l-1}^{-1} 
\]

\[
\mathcal{G}_{ll} = \sigma_l^2 \sqrt{\frac{N_{l-1}}{N_l}} G_{l-1,l}^{-1} \gamma_l.
\]

We see that \((28), (31)\) and \((32)\) form a closed set of 3L equations for 3L unknowns, \(\mathcal{G}_{ll}, \mathcal{G}_{ll} \) and \(\gamma_l\). Once solved, when the \(\gamma_l\)’s have been found, we are able to recover the Green’s function of \(\tilde{\mathbf{P}}\), and subsequently the \(M\)–transforms of \(\tilde{\mathbf{P}}\) and \(\mathbf{P}\) (in the argument \(w^L\)) \((13)\),

\[
\mathcal{G}_{ll} = \sigma_l^2 \sqrt{\frac{N_{l+1}}{N_l}} G_{l+1,l+1} G_{l,l-1}^{-1} 
\]

\[
\mathcal{G}_{ll} = \sigma_l^2 \sqrt{\frac{N_{l-1}}{N_l}} G_{l-1,l}^{-1} \gamma_l.
\]

where we have traded the \(\gamma_l\)’s for a more convenient set of variables,

\[
\mu_l \equiv |w|^2 \gamma_l - 1.
\]

Equations \((31)\) and \((32)\) form a set of decoupled recurrence relations for \(\mathcal{G}_{ll}\) and \(\mathcal{G}_{ll}\), respectively. Iterating these recurrences down to \(l = 1\) gives us:

\[
\mathcal{G}_{ll} = \mathcal{G}_l \left( \frac{1}{\sigma_1 \sigma_2 \ldots \sigma_{l-1}} \right)^2 \frac{1}{\sqrt{R_l}} \gamma_2 \gamma_2 \ldots \gamma_{l-1}.
\]

\[
G(w, \bar{w}) = \frac{1}{N_{\text{tot}} \sum_{i=1}^L N_i \gamma_i}, \quad i.e., \quad M(w, \bar{w}) = \frac{1}{N_{\text{tot}} \sum_{i=1}^L N_i \mu_i}, \quad i.e., \quad M(w^L, \bar{w}^L) = \frac{1}{L \sum_{i=1}^L R_i \mu_i},
\]

where we have traded the \(\gamma_i\)’s for a more convenient set of variables,

\[
\mu_l \equiv |w|^2 \gamma_l - 1.
\]
Applying the cyclic convention \(0 = L\), we get the following equation:

\[
\left(\sigma_1 \sigma_2 \ldots \sigma_L\right)^2 G_{\hat{\mathcal{T}}_1} = G_{\mathcal{T}}.
\] (37)

Straightforwardly, we get a trivial solution: \(G_l = 0\) for all \(l\), i.e., remembering (28), \(\gamma_l = 1/|w|^2\), or equivalently \(\mu_l = 0\) from (34), and therefore \(M_P(z, \bar{z}) = 0\) from (33). This is the holomorphic solution, holding outside the eigenvalue distribution. In order to retrieve information on the eigenvalue distribution, let us take \(G_{\hat{\mathcal{T}}_1} \neq 0\).

After a change of variables to \(\mu_l\) and some simplifications, (28) becomes

\[
\mu_l = \frac{|w|^2 G_{\hat{\mathcal{T}}_1} G_{\mathcal{T}_1}}{\mu_1 + 1} \frac{1}{R_l},
\] (38)

from which we get the relation:

\[
\frac{\mu_l}{R_l} = \frac{\mu_1}{R_l}.
\] (39)

After plugging (39) into (37), in terms of the \(\mu_l\) variables we obtain:

\[
(\mu_1 + 1) \left(\frac{\mu_1}{R_2} + 1\right) \ldots \left(\frac{\mu_1}{R_L} + 1\right) = \frac{|w|^2}{\sigma^2}.
\] (40)

On the other hand, substituting (39) into (33), we get

\[
M_P (\bar{w}_L, \bar{w}_L) = \mu_1.
\] (41)

All in all, after changing the argument from \(w\) to \(z = w_L\) we see that \(M_P(z, \bar{z})\) obeys the \(L\)-th order polynomial equation

\[
\left(\frac{M_P(z, \bar{z})}{R_1} + 1\right) \left(\frac{M_P(z, \bar{z})}{R_2} + 1\right) \ldots \left(\frac{M_P(z, \bar{z})}{R_L} + 1\right) = \frac{|z|^2}{\sigma^2},
\] (42)

which is precisely the first main result of our article, (9).

The last point to be addressed is to determine the validity domain of the non–holomorphic solution (42), knowing that on the boundary of such a domain, the non–holomorphic and holomorphic solutions must be joined. Thus, plugging the latter \((M_P(z, \bar{z}) = 0)\) into (42), we obtain an equation for the borderline:

\[
|z| = \sigma.
\] (43)

This means that the eigenvalues of the \(P\) matrix are scattered on average, with the density stemming from (42), within a centered circle of radius \(\sigma\).

When \(L = 2\), (42) is just a second degree equation, and it is easily solved. Indeed, in this case the non–holomorphic \(M\)–transform reads

\[
M_P(z, \bar{z}) = \frac{1}{2} \left(-1 - R + \sqrt{(1 - R)^2 + 4R|z|^2/\sigma^2}\right),
\] (44)

where we pose \(R \equiv R_2 = N_2/N_1\), and where the proper solution of (42) has been picked up in order to satisfy the matching condition (43) with the holomorphic one on the borderline. As a result, we immediately obtain the Green’s function:

\[
G_P(z, \bar{z}) = \frac{1}{2z} \left(1 - R + \sqrt{(1 - R)^2 + 4R|z|^2/\sigma^2}\right).
\] (45)
When deriving the average spectral density, one has to be cautious in the vicinity of the origin of the complex plane in order to properly take possible zero modes into account. Let us first expand near \( z = 0 \) in order to clarify its behavior:

\[
G_{P}(z, \overline{z}) \sim \frac{f}{z} + \text{regular terms}, \quad \text{as} \quad z \to 0, \quad \text{where} \quad f \equiv \begin{cases} 1 - R, & \text{for} \quad R < 1, \\ 0, & \text{for} \quad R \geq 1. \end{cases} \tag{46}
\]

Taking the derivative \((1/\pi)\partial_{z}\) of this singular term yields a Dirac delta function at the origin, \( f\delta^{(2)}(z, \overline{z}) \). Altogether,

\[
\rho_{P}(z, \overline{z}) = \begin{cases} \frac{1}{\pi\sigma^{2}} \frac{R}{\sqrt{(1 - R)^{2} + 4R|z|^{2}R}} + f\delta^{(2)}(z, \overline{z}), & \text{for} \quad |z| \leq \sigma, \\ 0, & \text{for} \quad |z| > \sigma. \end{cases} \tag{47}
\]

Moreover, one can also verify that the density, in the thermodynamic limit, changes on the borderline from being non-holomorphic with value given by

\[
\rho_{P}(z, \overline{z}) \big|_{|z| = \sigma} = \frac{1}{\pi\sigma^{2}} R_{h}, \quad \text{where} \quad \frac{1}{R_{h}} \equiv \sum_{i=1}^{L} \frac{1}{R_{i}},
\]

to being holomorphic with value 0. However, for finite sizes of the random matrices, this step gets smoothed out. Let us then consider the radial density,

\[
\rho_{P}^{\text{rad}}(r) \equiv 2\pi r \rho_{P}(z, \overline{z}) \big|_{|z| = r}
\]

and, following [16], let us propose the following model for this finite-\(N\) effect (where by \(N\) we denote the order of magnitude of the dimensions of the matrices, say \(N = N_{1}\)). We introduce an “effective” radial density in order to properly incorporate such finite-\(N\) behavior at the borderline,

\[
\rho_{P}^{\text{eff.}}(r) \equiv \frac{\rho_{P}^{\text{rad}}(r)}{2} \text{erfc} \left( q(r - \sigma)\sqrt{N} \right),
\]

where \(q\) is a free parameter whose value is to be adjusted by fitting. We numerically verify this hypothesis (see figures [1] and [2]).

**V. THE SINGULAR VALUES OF A PRODUCT OF RECTANGULAR GAUSSIAN RANDOM MATRICES**

In the following we show how to derive formula [10], i.e., an \((L + 1)\)-th order polynomial equation obeyed by the \(M\)-transform (which, as already discussed, encodes the same information contained in the spectral density) of the Hermitian matrix \(Q \equiv P^{\dagger}P\) [3], \(P\) being the product [11] of rectangular [2] Gaussian random matrices [3]. Our result agrees with that in [17], obtained in the context of wireless telecommunication theory, provided we synchronize the conventions. In particular, our resolvent \(G(z)\) relates to Stieltjes transform as \(G(z) = -G(-s)\). The underlying idea will be to rewrite \(Q\) as a product of some Hermitian matrices in order to apply the techniques provided by Free Random Variables (FRV) calculus. Loosely speaking, FRV calculus (initiated by the pioneering works of Speicher and Voiculescu et al.) can be thought as the extension of standard probability theory to non-commutative objects, such as matrices. Given the broadness of the topic, we shall not attempt any introductory discussion here, and we refer the non-expert reader to [12] [13].

Let us commence by defining, for any \(l = 1, 2, \ldots, L\), a square \(N_{l+1} \times N_{l+1}\) matrix

\[
Q_{l} \equiv (A_{1}A_{2} \ldots A_{l-1}A_{l})^{\dagger} (A_{1}A_{2} \ldots A_{l-1}A_{l}) = A_{l}^{\dagger}A_{l-1}^{\dagger} \ldots A_{2}^{\dagger}A_{1}^{\dagger}A_{1}A_{2} \ldots A_{l-1}A_{l},
\]

being a generalization of \(Q\) which includes only the first \(l\) random matrices, as well as a square \(N_{l} \times N_{l}\) matrix, which differs from \(Q_{l}\) only in the position of the last matrix in the string, i.e., \(A_{l}\), which is now placed as the first matrix in the string,

\[
\bar{Q}_{l} \equiv A_{1}A_{l}^{\dagger}A_{l-1}^{\dagger} \ldots A_{2}^{\dagger}A_{1}^{\dagger}A_{2} \ldots A_{l-1}A_{l} = (A_{1}A_{l}^{\dagger}) Q_{l-1}.
\]
FIG. 1: Numerical verification of the theoretical formula (4.7) for (the radial part (4.9) of) the mean spectral density \( \rho_P(z, \tau) \) of the product \( P \) of \( L = 2 \) rectangular Gaussian random matrices, as well as the finite–size correction (5.0).

UP LEFT: A numerical histogram (the black line) versus the theoretical prediction (4.7), supplemented with the finite–size smoothing (5.0) (the red plot), for \( N_1 = 100 \) and \( N_2 = 200 \) (i.e., \( R = R_2 = 2 \)), and for \( 10^5 \) Monte–Carlo iterations (i.e., the histogram is made of \( 10^7 \) eigenvalues). The adjustable parameter \( q \) (5.0) is fitted to be \( q \approx 1.14 \).

UP RIGHT: An analogous graph to UP LEFT, this time with \( N_1 = 100 \) and \( N_2 = 150 \) (i.e., \( R = 1.5 \)). We find \( q \approx 1.08 \) here.

DOWN LEFT: An analysis of the finite–size effects: numerical histograms for \( N_1 = 50, N_2 = 100 \) (black), \( N_1 = 100, N_2 = 200 \) (dashed red), \( N_1 = 200, N_2 = 400 \) (dotted blue), i.e., with the same rectangularity ratio \( R = 2 \), but increasing matrix dimensions. We observe how these plots approach the green line of the theoretical formula (4.7) for the density in the thermodynamic limit.

DOWN RIGHT: Numerical histograms for the matrix sizes of \( N_1 = 100, N_2 = 200 \) (i.e., \( R = 2 \); black) and \( N_1 = 200, N_2 = 100 \) (i.e., \( R = 1/2 \); red). Due to the presence of the zero modes (not displayed in the picture), the latter is half of the former.

We are interested in the eigenvalues of the Hermitian matrix \( Q = Q_L \).

The orders of the terms in the two above products (51), (52) are related to each other by a cyclic shift, therefore, for any integer \( n \geq 1 \), there will be \( \text{Tr} Q^n = \text{Tr} Q_n^\tau \). Hence, the \( M \)–transforms (see equation (6)) of the two above random matrices are related by the following relation

\[
M_{Q_1}(z) = \sum_{n \geq 1} \frac{1}{z^{n+1}} N_{l+1} \left\langle \text{Tr} Q_1^n \right\rangle = \frac{N_l}{N_{l+1}} \sum_{n \geq 1} \frac{1}{z^n} N_l \left\langle \text{Tr} Q_1^n \right\rangle = \frac{R_l}{R_{l+1}} M_{Q_1}(z). \tag{53}
\]

Now, let us consider the functional inverse of the \( M \)–transform, called the \( N \)–transform, defined as: \( M_{Q_1}(N_{Q_1}(z)) = N_{Q_1}(M_{Q_1}(z)) = z \). Employing this definition within equation (53) one easily obtains

\[
N_{Q_1}(z) = N_{Q_1} \left( \frac{R_{l+1}}{R_l} z \right). \tag{54}
\]
Following recurrence relation for the \( N \) with the initial condition, \( \text{applied. Such law states that the } N \text{ is a product of two free matrices, } A_i A_j^\dagger \text{ and } Q_{i-1}. \) Then, the FRV multiplication law for free matrices can be applied. Such law states that the \( N \)-transform of the product of two free matrices, \( A \) and \( B \), is simply given by \( N_{AB}(z) = z/(1+z) N_A(z)N_B(z). \) (In the language more often found in the literature on the subject, the \( N \)-transform is replaced by the so-called \( S \)-transform, \( S_X(z) \equiv (z+1)/(zN_X(z)) \), which then obeys a simpler multiplication law, \( S_{AB}(z) = S_A(z)S_B(z) \).) So, when applying this relation to the \( Q \) matrix \( [32] \) one can write, for \( l = 2, 3, \ldots, L \)

\[
N_{Q_l}(z) = \frac{z}{z+1} N_{A_i A_j^\dagger}(z) N_{Q_{l-1}}(z). \tag{55}
\]

From equations \([54]\) and \([55]\), we now eliminate the \( N \)-transform of the auxiliary \( \tilde{Q}_l \), which leaves us with the following recurrence relation for the \( N \)-transform of \( Q_l \),

\[
N_{Q_l}(z) = \frac{z}{z+1} N_{A_i A_j^\dagger}(z) N_{Q_{l-1}}(z), \quad \text{for} \quad l = 2, 3, \ldots, L, \tag{56}
\]

with the initial condition,

\[
N_{Q_1}(z) = N_{\tilde{Q}_1} \left( \frac{R_2}{R_1} \frac{z}{R_1} \right) = N_{A_i A_j^\dagger} \left( \frac{R_2}{R_1} \frac{z}{R_1} \right), \tag{57}
\]

which stems from \([32]\) and from the fact that \( \tilde{Q}_1 = A_i A_j^\dagger \). The solution of this recurrence \([56], [57]\) is then readily found to be

\[
N_{Q_L}(z) = \frac{z^{L-1}}{(z+R_2)(z+R_3)\ldots(z+R_L)} N_{A_i A_j^\dagger} \left( \frac{z}{R_1} \right) N_{A_j A_k^\dagger} \left( \frac{z}{R_2} \right) \ldots N_{A_{L-1} A_L^\dagger} \left( \frac{z}{R_L} \right). \tag{58}
\]

It remains now to find the \( N \)-transforms of the random matrices \( A_i A_j^\dagger \). They are examples of the so-called “Wishart ensembles”, and the problem of computing their \( N \)-transforms, with the same normalization of the probability measures \([3]\) of the \( A_i \)'s which we are employing, has first been solved in \([3]\): expressions (1.8), (2.8), (2.13), (2.14) of this article yield the Green’s function of \( A_i A_j^\dagger \), which immediately leads to the pertinent \( N \)-transform,

\[
N_{A_i A_j^\dagger}(z) = \sigma_i^2 \left( \frac{z+1}{\sqrt{N_{i+1}z+\sqrt{N_i}}} \right). \tag{59}
\]

Substituting \([59]\) into \([58]\), one finally arrives at the desired formula for the \( N \)-transform of \( Q = Q_L \),

\[
N_Q(z) = \sigma^2 \sqrt{R_1} \left( \frac{z+1}{R_1+1} \right) \left( \frac{z}{R_2+1} \right) \ldots \left( \frac{z}{R_L+1} \right), \tag{60}
\]

FIG. 2: Analogous graphs to figure [1] DOWN LEFT, but for \( L = 3 \) (LEFT) and \( L = 4 \) (RIGHT).
with \( \sigma \) defined as in the previous sections. In other words, the corresponding \( M \)-transform \( M_Q(z) \) satisfies the following polynomial equation of order \((L + 1)\),

\[
\sqrt{R_1} \frac{1}{M_Q(z)} (M_Q(z) + 1) \left( \frac{M_Q(z)}{R_1} + 1 \right) \left( \frac{M_Q(z)}{R_2} + 1 \right) \ldots \left( \frac{M_Q(z)}{R_L} + 1 \right) = \frac{z}{\sigma^2},
\]

(61)

or in the case of \( N_{L+1} = N_1 \) (i.e., \( R_1 = 1 \), required when one wishes for \( P \) to have eigenvalues too),

\[
\frac{1}{M_Q(z)} (M_Q(z) + 1)^2 \left( \frac{M_Q(z)}{R_2} + 1 \right) \ldots \left( \frac{M_Q(z)}{R_L} + 1 \right) = \frac{z}{\sigma^2}.
\]

(62)

This completes our derivation of (11).

![Figure 3: Numerical verification of the theoretical formula (61) for the mean spectral density](image)

FIG. 3: Numerical verification of the theoretical formula (61) for the mean spectral density \( \rho_Q(\lambda) \) of the random matrix \( Q = P^T P \). Everywhere we have \( N_{L+1} = N_1 = 50 \). The number of Monte–Carlo iterations is 20,000, i.e., all the histograms are generated from \( 10^6 \) eigenvalues.

LEFT: \( L = 2 \), and the matrix sizes are chosen to be \( N_1 = 50, N_2 = 150 \).

MIDDLE: \( L = 3 \), and the matrix sizes are \( N_1 = 50, N_2 = 100, N_3 = 150 \).

RIGHT: \( L = 4 \), and the matrix sizes are \( N_1 = 50, N_2 = 100, N_3 = 150, N_4 = 200 \).

We have performed extended numerical tests of the formula (61), in all cases obtaining perfect agreement, see figure 3.

VI. CONCLUSIONS

The main contribution of this article is equation (9) for the \( M \)-transforms of the product \( P = A_1 A_2 \ldots A_L \) of an arbitrary number \( L \) of independent rectangular \( \mathbb{C} \) Gaussian random matrices \( A_i \). Knowing the \( M \)-transform one can easily calculate the eigenvalue density of the product (7), which turns out to be spherically symmetric in the complex plane. We also discussed a striking resemblance of equation (9) to the corresponding equation (10) of the Hermitian matrix \( Q = P^T P \), whose eigenvalues are equal to the squared singular values of \( P \). Both these equations are polynomial (of orders \( L \) and \((L + 1)\) respectively), so in general they may only be solved numerically; however, some properties of the mean spectral densities can still be retrieved analytically, such as their singular behavior at zero \( \lambda \).

We are tempted to conjecture that this similarity of the \( M \)-transforms for \( P \) and \( Q \) is generic for random matrices possessing rotationally symmetric average distribution of the eigenvalues, and that the corresponding equations differ only by the prefactor which we have discussed while comparing (9) and (10). For such models, the non–holomorphic \( M \)-transform \( M_X(z, \bar{z}) \) is a function of the real argument \(|z|^2\), thereby allowing for functional inversion, and hence for a definition of the “rotationally–symmetric non–holomorphic \( N \)-transform” — even though for general non–Hermitian random matrices a construction of a “non–holomorphic \( N \)-transform” remains thus far unknown. This new \( N \)-transform is then conjectured to be in a simple relation to the (usual) \( N \)-transform of the Hermitian ensemble \( X \). In a typical situation, the latter will be much more easily solvable than the former, owing to the plethora of tools devised in the Hermitian world, albeit the opposite may be true as well. This is indeed the case here — our derivation of (9), based on non–Hermitian planar diagrammatics and Dyson–Schwinger’s equations, is much more involved than a simple application of the FRV multiplication rule leading to (10) — and consequently, the aforementioned hypothesis would provide a shortcut to avoid complicated diagrammatics. To the best of our knowledge, this would be the first
use of Free Random Variables calculus to compute the mean spectral density of a non–Hermitian product of random matrices.

We have also suggested a heuristic model of the finite–size behavior of the density of \( P \) near the edge of the eigenvalues support \([59]\), deducing it from analogous considerations \([31, 32]\) made for the Girko–Ginibre ensemble, where this behavior is known analytically. It performs outstandingly well when checked against numerical simulations.

Let us also remark that one could argue, as for square matrices, that the large– \( N \) limit result is the same for elliptic Gaussian ensembles \([10]\). We also believe that one can further weaken the assumptions on the matrices involved, just requiring them to belong to the Gaussian universality class of matrices having independent entries and fulfilling the Pastur–Lindeberg condition \([38]\) (the matrix analogue of the generalized central limit theorem in classical probability theory \([39]\)). One unexpected implication of such universality is that a product of random matrices whose spectra do not necessarily display rotational symmetry has an eigenvalue distribution which does possess rotational symmetry on the complex plane \( i.e., \) the average density depends only on \(|\lambda|\)).

Let us now list some possible applications of these results to wireless telecommunication, quantum entanglement and multivariate statistical analysis.

Information theory for wireless telecommunication has been intensively developed in the past decade, after it had been realized that in a number of situations the information transmission rate can be increased by an introduction of multiple antenna channels, known as the “multiple–input, multiple–output” (MIMO) transmission links. The MIMO capacity for Gaussian channels has been calculated in the pioneering work \([35]\), triggering large activity in the field. Immediately, it became clear that an appropriate language and methods to address this type of problems are provided by random matrix theory (consult \([5]\) for a review). The model considered in our paper can be applied to a situation of signals traveling over \( L \) consecutive MIMO links. The signal is first sent from \( N_1 \) transmitters via a MIMO link to \( N_2 \) receivers, which then re–transmit it via a new MIMO link to the subsequent \( N_3 \) receivers, \( etc. \) Clearly, the capacity will depend on these numbers of intermediate re–transmitters; in particular, if any of the \( N_i \)'s is small, the capacity will be reduced. The effective propagation is given by the matrix \( P = A_L \ldots A_2 A_1 \). Such a model of multifold scattering per propagation path has been already proposed in \([17]\), where the moment generating function, the \( M \)–transform, for \( P^\dagger P \) was calculated. Our result for the \( M \)–transform for \( P \) complements this calculation.

Let us also mention that one could imagine a more general situation, where MIMO links form a directed network — each directed link \( lm \) representing a single MIMO channel between \( N_l \) transmitters and \( N_m \) receivers. (The previously discussed case corresponds to a linear graph, \( 1 \rightarrow 2 \rightarrow \ldots \rightarrow L \).) A complex directed network of MIMO links is somewhat similar to the structures appearing in the context of quantum entanglement. There, one considers graphs whose edges describe bi–partite maximally entangled states, while vertices describe the couplings between subsystems residing at the same vertex \([39]\). In the simplest case of a graph consisting of a single link, it is just a bi–partite entangled state. The corresponding density matrix for a bi–partite subsystem is given by \( Q = A^\dagger A \), where \( A \) is a rectangular matrix defining a pure state, being a combination of the basis states in the subsystem, \( |\alpha_a\rangle \) and \( |\beta_b\rangle \) (see for instance \([20]\)). One can easily find that linear graphs with additional loops at the end vertices correspond. The density matrix for the subsystem sitting in the end vertex is given by \( Q = P^\dagger P \), where \( P = A_1 A_2 \ldots A_L \) \([30]\). If all the subsystems are of the same size, the average spectral distributions \([14, 15]\) of \( Q \) are known as the “Fuss–Catalan family” \([37]\); they can be obtained from \([10]\) by setting all the \( R_i \)'s to 1. However, if the subsystems have different sizes, one needs to apply our general formula \([10]\).

Finally, another area of applications of our approach is related to multivariate analysis. The main building block there is the Wishart ensemble, corresponding to \( L = 1 \) in our formalism. The link between the spectral properties of \( P \) and \( Q \) may allow one to avoid the well–known bottleneck caused by the non–Hermiticity of time–lagging correlation functions. This issue will be discussed in a forthcoming publication.

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