 Associations between community well-being and hospitalisation rates: results from a cross-sectional study within six US states

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ABSTRACT

Objective To evaluate the association between community well-being, a positively framed, multidimensional assessment of the health and quality of life of a geographic community, and hospitalisation rates.

Design Cross-sectional study

Setting Zip codes within six US states (Florida, Iowa, Nebraska, New York, Pennsylvania and Utah)

Main outcome measures Our primary outcome was age-adjusted, all-cause hospitalisation rates in 2010; secondary outcomes included potentially preventable disease-specific hospitalisation rates, including cardiovascular-related, respiratory-related and cancer-related admissions. Our main independent variable was the Gallup-Sharecare Well-Being Index (WBI) and its domains (life evaluation, emotional health, work environment, physical health, healthy behaviours and basic access).

Results Zip codes with the highest quintile of well-being had 223 fewer hospitalisations per 100 000 (100k) residents than zip codes with the lowest well-being. In our final model, adjusted for WBI respondent age, sex, race/ethnicity and income, and zip code number of hospital beds, primary care physician density, hospital density and admission rates for two low-variation conditions, a 1 SD increase in WBI was associated with 5 fewer admissions/100k (95% CI 4.0 to 5.8; p<0.001). Results were similar for cardiovascular-related and respiratory-related admissions, but no association remained for cancer-related hospitalisation rates. Patterns were similar for each of the WBI domains and all-cause hospitalisations.

Conclusion and relevance Community well-being is inversely associated with local hospitalisation rates. In addition to health and quality-of-life benefits, higher community well-being may also result in fewer unnecessary hospitalisations.

INTRODUCTION

Hospitalisation rates vary widely across the USA, with county-level hospitalisation rates for ambulatory sensitive conditions ranging from 14 to 281 hospitalisations per 1000 Medicare enrollees in 2011. This 20-fold variation in potentially preventable hospitalisation rates by location suggests that characteristics of the community in which patients live may contribute to higher or lower hospitalisation rates.

As part of efforts to reduce unnecessary hospitalisation rates, health systems and other stakeholders have begun to seek modifiable community properties that are associated with lower hospitalisation rates, and may, therefore, serve as potential targets for intervention. Prior studies have identified some community factors that are linked to rates of hospitalisations, including housing instability and access to primary care, as well as road and airport traffic noise, air quality and smoking policies.
Emerging evidence suggests that psychosocial factors may also be linked to hospitalisation rates. At the individual level, a broad measure of subjective well-being—a comprehensive measure comprised of physical and emotional health, healthy behaviours, life evaluation, work environment and access to basic needs—was associated with fewer emergency department visits, hospitalisations and healthcare costs.\textsuperscript{8, 10} It is plausible that this inverse association with well-being extends to the community level. Our team has previously linked higher community-level positive psychosocial factors, such as social capital, have been reported to be protective of cardiovascular hospitalisations in prospective studies. However, the relationship between population well-being and hospitalisation risk has yet to be examined. Therefore, we sought to evaluate a community-level association between well-being and hospitalisation, hypothesising that higher community well-being is associated with lower rates of hospitalisation.

**METHODS**

**Overview**

We linked data from a comprehensive, multidimensional assessment of well-being, the Gallup-Sharecare Well-Being Index (WBI), with data on hospitalisation rates from a diverse set of six states from state all-payer claims databases or the Agency for Healthcare Research and Quality’s Healthcare Costs and Utilization Project (HCUP). All data were from the year 2010, as this was the most recent year for which we had access to HCUP data and the year that had the largest sample size for our covariates. All data were aggregated and linked at the zip code level. We performed a cross-sectional study to assess the correlation between community well-being and hospitalisation rates.

**Sample**

We included data on all adult and paediatric hospitalisations from zip codes within six states: Florida, Iowa, Nebraska, New York, Pennsylvania and Utah. We selected states with available data on hospitalisation rates by zip code for the calendar year 2010 through state all-payer claims databases or the HCUP. Within these restrictions, we aimed to choose a diverse sample of states from different regions of the USA and with urban, suburban and rural areas. We excluded hospitalisations for patients with listed home zip codes outside the state where they were hospitalised. We also excluded zip codes with fewer than 10 survey participants in the WBI in the year 2010.

**Community well-being**

Our main independent variable was the composite well-being score, as measured by the Gallup-Sharecare WBI. We used WBI data from 01 January 2010 to 31 December 2010, aggregated by zip code, to assess community well-being.

The Gallup-Sharecare WBI is an ongoing telephone-based survey administered to a random sample of 500–1000 adults in the USA daily, 350 days per year, since 2008. The random sample includes adults aged 18 years and older residing in the USA, who speak either English or Spanish and have either a landline or cellular phone. Gallup used a structured sampling design to obtain representative data from all 50 states and the District of Columbia.\textsuperscript{12}

The WBI was developed based on the work of psychology experts.\textsuperscript{13, 14} Briefly, survey items that aligned with previous research on well-being were initially compiled by experts in the field.\textsuperscript{15, 16} Based on reviews of the literature, items were selected to encompass both hedonic well-being (ie, people’s feelings and thoughts about their lives)\textsuperscript{17} and eudemonic well-being (ie, an individual’s judgments about the meaning and purpose in one’s life),\textsuperscript{18} and thus incorporated items assessing daily emotional experience and a wide variety of evaluative domains, such as overall life, standard of living, and satisfaction with community, work, relationships and personal health. Factor analysis using data from a large, representative national sample was then used to determine the final set of questions. Criterion validity of regionally aggregated data was established by examining correlations with regional health and socioeconomic indicators.\textsuperscript{19} Subsequently, principal component analyses and confirmatory factor analyses were used to create an instrument valid for measuring well-being at the individual level. The individual well-being measure has acceptable reliability, and internal and external validity.\textsuperscript{20}

The WBI includes 40 self-reported items organised into six domains representing key aspects of well-being,\textsuperscript{15} life evaluation (life satisfaction and optimism), emotional health (positive and negative affect and the presence or absence of depression), work environment (job satisfaction and trust and respect in the workplace among respondents who report being employed), physical health (chronic disease and recent illness), healthy behaviours (smoking, exercise, and fruit and vegetable consumption) and basic access (perceived safety, housing and healthcare access).\textsuperscript{12} Each domain is measured on a scale of 0–100, and the composite well-being score is an unweighted mean of all the six domains.\textsuperscript{21}

**Hospitalisation rates**

Our dependent variable was age-adjusted, all-cause hospitalisation rate. Hospitalisation rates by patient’s residence zip code were obtained using all-payer claims databases or data from the HCUP from Florida, Iowa, Nebraska, New York, Pennsylvania and Utah. Patients with out-of-state residence zip codes were excluded. Rates were expressed in number per 100000 (100k).

To guide future work on reducing potentially preventable hospitalisations, we also examined rates of cardiovascular-related, respiratory-related and specific cancer-related hospitalisations as additional dependent variables. Major cardiovascular disease (CVD) discharges
(eg, ischaemic heart disease, hypertensive disease, rheumatic heart disease and stroke) were identified by International Classification of Disease, Ninth Revision, Clinical Modification (ICD-9-CM) codes of 390–448 as the primary diagnosis. Respiratory-related hospitalisations included hospital discharges for asthma, chronic obstructive pulmonary disease and pneumonia (ICD-9-CM primary diagnosis codes 480–496). Specific cancers included in this study were those with standard recommended screening protocols, including breast cancer, lung cancer, colon cancer and cervical cancer (ICD-9-CM primary discharge diagnosis codes 153–4, 162, 174 and 180).

**Covariates**

Covariates included zip code population estimates for age, sex, race/ethnicity and median income, obtained from the 2010 US Census. We also included the following healthcare intensity variables obtained from the Area Health Resource File22: number of hospital beds, primary care physician density and hospital density. Hospital beds were defined as the number of beds regularly maintained (set up and staffed for use) for inpatients as of the close of the reporting period. Primary care physician density was defined as non-Federal doctors of medicine and doctors of osteopathy providing direct patient care who practice principally in one of the four primary care specialties—general or family practice, general internal medicine, paediatrics, and obstetrics and gynaecology per 100k population. Hospital density was defined as the number of acute care hospitals per 100k population. Finally, to account for differences in hospital referral regions and volume, we included admission rates for two low-variation conditions for hospitalisation not included as a secondary outcome (hip fracture and gastrointestinal bleeding) from the state all-payer claims databases or data from the HCUP.

**Analysis**

We first described the distribution of well-being scores and prevalence of healthcare intensity variables across included zip codes, stratified by quintile of hospitalisation rates. Next, we described the mean all-cause, cardiovascular, respiratory and cancer-related hospitalisation rates by quintile of composite well-being score and by each well-being domain.

We then estimated three nested multivariable linear regression models to explore associations between 1 SD increase in the composite well-being score and each well-being domain with zip code hospitalisation rates. We first estimated an unadjusted model (model 1) which included only indicators for state, and sequentially adjusted for additional sets of covariates. A second model (model 2) adjusts for population age, sex, race and income, and admission rates for low-variation conditions. The final model (model 3) further adjusts model 2 for healthcare intensity variables. We repeated these analyses for cardiovascular-related, respiratory-related, and cancer-related hospitalisations. We used the method of Belsley, Kuh and Welsch to assess the covariates for collinearity; using their criteria, we excluded the proportion of respondents aged between 65 and 84 years from the models.24

All analyses were performed using Stata V.15.0.

**Patient and public involvement**

No patients or the public was involved in the planning and design of this study.

**RESULTS**

We included 1487 zip codes in our analysis. The mean WBI score ranged from 56 in the lowest quintile up to 78.2 in the highest quintile (table 1). Zip codes with higher WBI tended to have fewer adults aged over 65 years, though there was no difference in any other age groups. Zip codes with higher WBI also had higher median household income, but there was no difference in per cent white residents. In addition, zip codes with higher WBI had more general practitioners per 100k residents. No other consistent patterns of healthcare intensity variables were noted across quintiles of well-being.

As the zip code WBI increased from the lowest to highest quintile, the mean number of hospitalisations per 100k residents decreased (figure 1). Zip codes in the highest quintile of well-being had 17% fewer hospitalisations per 100k residents compared to the lowest quintile. This finding was consistent across all causes of hospitalisation, with the largest reductions observed for cancer-related hospitalisations (17.5% lower in the highest quintile compared to the lowest quintile).

### Table 1 Differences in mean sociodemographic factors and healthcare intensity variables across quintiles of well-being *(n=1487 zip codes)*

|                | Q1     | Q2     | Q3     | Q4     | Q5     | P trend |
|----------------|--------|--------|--------|--------|--------|---------|
| **Well-being** | 56.0   | 64.0   | 67.9   | 71.5   | 78.2   | N/A     |
| % age>65 years | 11.1   | 11.2   | 10.5   | 10.7   | 10.5   | 0.02    |
| % white       | 86.9   | 89.2   | 89.3   | 90.1   | 91.5   | 0.62    |
| Median household income (thousands) | 48.0   | 51.1   | 57.6   | 60.4   | 64.5   | <0.001  |
| PCP/100k      | 27.1   | 28.6   | 31.5   | 31.4   | 31.0   | <0.001  |
| Hospitals/100k | 2.9    | 2.2    | 2.4    | 2.8    | 3.1    | 0.23    |
| Hospital beds/100k | 0.3    | 0.3    | 0.4    | 0.3    | 0.3    | 0.38    |

100k, 100 000; N/A, not available; PCP, primary care provider; Q, quintile.
hospitalisations than zip codes in the lowest quintile of well-being (an absolute difference of 223 per 100k residents). After adjustment for age, sex, race, income and healthcare intensity variables, zip codes in the highest quintile of well-being had 9% fewer hospitalisations than zip codes in the lowest quintile of well-being (absolute difference of 12 per 100k residents; online supplementary appendix). The number of hospitalisations per 100k decreased in a stepwise fashion across quintiles of well-being. Similar patterns of smaller magnitude were seen for cardiovascular-related (absolute difference: 35/100k; 20% fewer hospitalisations), respiratory-related (absolute difference: 19/100k; 26% fewer hospitalisations) and cancer-related hospitalisations (absolute difference: 2/100k; 17% fewer hospitalisations).

In our unadjusted model, a 1 SD higher WBI was associated with 7.3 fewer all-cause hospital admissions/100k (p<0.001) (table 2). After adjustment for demographics and low-variation causes for hospital admission, a 1 SD higher WBI was associated with 5.5 fewer hospital admissions/100k. In our fully adjusted model, a 1 SD higher WBI was associated with 5 fewer all-cause hospital admissions/100k (p<0.001). Results were similar for cardiovascular-related and respiratory-related admissions, but no association remained for cancer-related hospitalisation after adjustment. The R² values for the fully adjusted WBI model ranged from 0.13 for cancer admissions to 0.43 for CVD admissions; the model for all-cause admissions had an R²=0.38 (table 3).

In general, similar patterns were seen for each of the well-being domains. All domains of well-being were inversely associated with all-cause hospitalisations in unadjusted and adjusted models. The strongest inverse relationships with all-cause hospital admissions were noted with a 1 SD increase in the basic access (14.7 fewer admissions/100k; p<0.001), physical health (6.6 fewer admissions/100k; p<0.001) and emotional health (6.6 fewer admissions/100k; p<0.001) domains in unadjusted models. In fully adjusted models, all-cause hospitalisations were most strongly associated with the basic access (8.3 fewer admissions/100k; p<0.001), physical health (4.9 fewer admissions/100k; p<0.001) and emotional health (4.6 fewer admissions/100k; p<0.001) domains.

**DISCUSSION**

Our study examines the association between a holistic, multidimensional assessment of well-being and hospitalisation rates by zip code. We found that higher community well-being was associated with lower all-cause, cardiovascular and respiratory hospitalisation rates, after controlling for sociodemographic and healthcare intensity variables, including density of primary care physicians. If all zip codes in the USA had well-being scores in the highest quintile, based on our findings, we would expect approximately 100k fewer hospitalisations yearly, saving nearly US$1 billion in healthcare costs. These results suggest that beyond access to care and social risk factors, there are additional community attributes that may influence local hospitalisation rates.

Community well-being is modifiable and supporting programmes and policies to foster community well-being should be a society’s end goal. Our results suggest that promoting community well-being may have an additional benefit of curbing unnecessary hospitalisation rates, as well as supporting other beneficial population health outcomes, including longer life expectancy and lower rates of preterm birth. Our team also recently reported that Medicare spent a mean of US$992 less per beneficiary who lived in counties with the highest quintile of well-being compared with those who lived in counties with the lowest quintile of well-being. Because hospitalisations are one of the largest drivers of healthcare costs, results from the present study support and expand our prior finding, in that the well-being of a community is linked to healthcare utilisation and costs across all ages and payers. In both studies, basic access and emotional health were among the three domains most strongly associated with the outcome. As hospitals and health systems seek to achieve lower total costs of care in value-based

![Figure 1](https://example.com/figure1.png)

**Figure 1** Median number and IQR of age-standardised and sex-standardised hospitalisations/100 000 (100k) population across quintiles of well-being (n=1487 zip codes). WBI, well-being index.

| Model 1 | Model 2 | Model 3 |
|---------|---------|---------|
| All-cause | −7.3** | −5.5** | −4.9** |
| Cardiovascular | −0.99** | −0.67** | −0.56** |
| Respiratory | −0.69** | −0.42** | −0.33** |
| Cancer | −0.05* | −0.04* | −0.04 |

Model 1: unadjusted; model 2: adjusted for sociodemographics and low-variation admissions; and model 3: further adjusted for healthcare intensity variables.

*p<0.05 and **p<0.001.
Table 3  Fully adjusted models assessing associations between zip code WBI and all-cause, cardiovascular-related (CVD), respiratory-related and cancer-related hospital admissions

|                | All-cause | CVD          | Respiratory | Cancer       |
|----------------|-----------|--------------|-------------|--------------|
|                | Coeff (SE)| Coeff (SE)   | Coeff (SE)  | Coeff (SE)   |
| WBI            | −48.55 (8.87) | −5.58 (1.40) | −3.26 (0.93) | −0.34 (0.19) |
| State          |           |              |             |              |
| FL             | ref       | ref          | ref         | ref          |
| IA             | −20.41 (4.26) | −4.23 (0.67) | −2.22 (0.45) | −0.11 (0.09) |
| NE             | −18.26 (4.44) | −4.67 (0.70) | −1.70 (0.47) | 0.03 (0.10)  |
| NY             | −20.82 (3.67) | −3.80 (0.58) | −2.01 (0.39) | 0.07 (0.08)  |
| PA             | −2.62 (3.26)  | −0.64 (0.51) | −1.25 (0.34) | 0.05 (0.07)  |
| UT             | −24.22 (3.93) | −7.73 (0.62) | −3.21 (0.41) | −0.36 (0.09) |
| % female       | 0.09 (0.04)  | 0.01 (0.01)  | 0.00 (0.00)  | 0.00 (0.00)  |
| Age, years     |           |              |             |              |
| % age 25–44    | 0.14 (0.06)  | 0.03 (0.01)  | 0.01 (0.01)  | 0.00 (0.00)  |
| % age 45–64    | −0.03 (0.04) | −0.00 (0.01) | 0.00 (0.01)  | 0.00 (0.00)  |
| % age 65–85    | −0.21 (0.11) | −0.03 (0.02) | −0.03 (0.01) | 0.00 (0.00)  |
| Income         |           |              |             |              |
| % 12–35k       | −0.02 (0.05) | 0.01 (0.01)  | 0.00 (0.01)  | −0.00 (0.00) |
| % 36–59k       | −0.04 (0.06) | 0.00 (0.01)  | −0.00 (0.01) | −0.00 (0.00) |
| % 60–120k      | −0.25 (0.06) | −0.03 (0.01) | −0.03 (0.01) | −0.01 (0.00) |
| % >120k        | −0.37 (0.08) | −0.06 (0.01) | −0.05 (0.01) | −0.01 (0.00) |
| Race           |           |              |             |              |
| % AA           | 0.84 (0.07)  | 0.10 (0.01)  | 0.06 (0.01)  | 0.01 (0.00)  |
| % Hispanic     | 0.75 (0.27)  | 0.05 (0.04)  | 0.04 (0.03)  | −0.00 (0.01) |
| % Asian        | −0.29 (0.28) | −0.11 (0.04) | −0.06 (0.03) | −0.01 (0.01) |
| % other        | 0.41 (0.12)  | 0.04 (0.02)  | 0.03 (0.01)  | −0.00 (0.00) |
| GPs/100k       | −2.81 (0.61) | −0.60 (0.10) | −0.44 (0.06) | −0.02 (0.01) |
| Hosp/100k      | −2.79 (0.94) | −0.40 (0.15) | 0.35 (0.10)  | −0.05 (0.02) |
| Hospital beds/100k | 7.44 (0.77) | 0.76 (0.12)  | 0.43 (0.08)  | 0.08 (0.02)  |
| Intercept      | 151.46 (8.52) | 21.21 (1.34) | 9.34 (0.90)  | 1.26 (0.19)  |
| $R^2$          | 0.39       | 0.43         | 0.29        | 0.13         |

AA, African–American; Coeff, coefficient; CVD, cardiovascular disease; FL, Florida; GPs, general practitioners; Hosp, hospital(s); IA, Iowa; k, thousands; NE, Nebraska; NY, New York; PA, Pennsylvania; ref, reference; UT, Utah; WBI, well-being index.

payment arrangements, paying attention to these community psychosocial attributes may foster success.

We report cross-sectional associations at the zip code level, and thus, we are unable to make causal inferences. However, it suggests that the efforts communities and policymakers are engaged in globally to improve community well-being may also have a modest effect on reducing hospitalisation rates. Our team recently reported that there are modifiable community characteristics, such as education levels, adequate income, diversity, rates of preventive care and rates of active commuting, that are associated with higher well-being of community residents. In addition, a prospective study in which municipal officials approved and permitted community member-driven environmental changes (eg, street murals, public benches and planter boxes, in public spaces) also resulted in improved well-being. Residents within a two-block radius of the transformed sites were systematically sampled and reported improvements in mental health, increased sense of community and an overall expansion of social capital. More comprehensive, longitudinal studies, such as the ongoing Well London Study, are needed to assess whether systemic community-based interventions influence well-being, and if they result in better population health outcomes, as hypothesised.

In general, these community-based interventions that improve well-being, such as improving housing conditions, income, safety and education, are linked to lower rates of hospitalisation. As such, promoting community well-being as a target outcome is likely to also result in
reducing unnecessary hospitalisations as community conditions are improved. However, our study suggests that there is a small but important additional reduction in hospitalisations attributable to higher community well-being itself. There are several potential mechanisms underlying this observed association. First, it is possible that the observed association is simply a result of aggregating individuals with higher well-being and their associated lower rate of hospitalisation. In two unique large employer-based populations, higher individual well-being predicted fewer hospitalisations and emergency department visits over the next 12 months. Similarly, among a large, multistate sample of Veterans, 10-point decrements in the physical and mental component scores of the Short Form-36 health-related quality of life were associated with higher rates of hospitalisations, after adjusting for sociodemographics, smoking and prior healthcare utilisation. Due to limitations in the data we had available, we were unable to test the relationship between individual and community well-being in this study, but even if the observed zip code-level association is due to the effects of higher aggregate individual resident well-being, it suggests that certain communities embody characteristics that foster higher resident well-being—a positive end in and of itself.

Community well-being may influence hospitalisation rates above and beyond the effect of individual well-being. This hypothesised extension of individual well-being to community well-being is plausible in that components of well-being, including happiness and altruism, that have been reported to diffuse across one’s social network in experimental and prospective observational studies. An increase in happiness of one person in a geographic community led to an increase in happiness among friends and family members who lived within one mile, but did not influence the happiness of co-workers, supporting the theory that higher well-being of individual residents in a community increases the well-being of the entire community. This logic is further supported by prior studies reporting other community-level positive psychosocial attributes are also directly associated with better health outcomes and inversely associated with healthcare utilisation. Our domain-specific analysis is aligned with these prior studies, as the emotional WBI alone was independently associated with fewer hospitalisations. Further, a systematic review of prospective, multilevel studies generally found direct relationships between positive community psychosocial attributes, such as social cohesion, social capital and collective efficacy, and health outcomes. Social capital was specifically protective of cardiovascular-related hospitalisations. After accounting for individual-level risk factors, residents living in neighbourhoods with low social capital had approximately 20% or 30% higher risk of first hospitalisation for fatal or non-fatal coronary disease, among men and women, respectively, over a 2-year period.

There are several potential mechanisms by which positive community psychosocial factors like well-being could exert an effect on hospitalisations: supporting healthy behaviours and social support, lowering the physiologic stress response and increasing perceived access to care. Community social cohesion and connectedness have been reported to be linked to the adoption of physical activity and not smoking, which influence the incidence of cardiovascular-related and respiratory-related hospitalisations. These associations persisted after adjustment for individual-level sociodemographics. These social factors are thought to influence health behaviours by promoting the spread of health information or increasing the likelihood that healthy norms of behaviour are adopted. In addition to supporting better health behaviours, stronger social support has been directly linked to fewer hospitalisations. Among older adults with heart failure, those with high levels of social support experienced fewer hospitalisations compared with those with moderate or low social support, in a graded fashion, over a 6.5-month follow-up period. Further, among multifactorial models predicting risk for hospital readmission, inclusion of social support variables improved overall model performance compared with using sociodemographic and clinical factors alone, supporting its unique and additive role in possibly preventing unnecessary hospitalisations.

Higher community well-being may also reduce hospitalisation rates by reducing or buffering the stress response. Chronic stress results in increased inflammation and higher allostatic load, which may reduce healing from low-acuity conditions and increase risk of CVD, thus resulting in higher rates of all-cause and cardiovascular-related hospital admissions, consistent with our findings. Most of the work supporting this physiologic pathway has been performed at the individual level, with higher levels of optimism and life satisfaction being associated with lower levels of chronic inflammation and incident CVD and mortality. But, because these positive psychologic factors propagate across social groups, these findings may extend to the community level as well, and emerging evidence suggests that higher levels of neighbourhood social well-being are indeed associated with lower levels of perceived chronic stress among residents. This logic is further supported by findings from a large population-based cohort study that reported that groups with the highest quintile of perceived stress experienced an approximately 1.5-fold higher risk of hospitalisation for ambulatory sensitive conditions, after adjusting for demographics, predisposing conditions, multimorbidity and socioeconomic factors.

Finally, perceived access to healthcare, beyond objective access to healthcare, is another mechanism by which communities with higher well-being may benefit. Higher well-being is associated with greater perceived access to healthcare, and prior studies report inverse associations between self-reported access to care and preventable hospitalisations in other US states and in countries with universal access to healthcare. In an Australian study, perceived access to healthcare was associated with
lower rates of admission for ambulatory sensitive conditions after controlling for propensity to seek healthcare, prevalence of ambulatory sensitive conditions, disease burden and the supply of primary care physicians.60 Our results strengthen these findings. Though we controlled for objective measures of access to healthcare, including density of primary care physicians, number of hospital beds, hospital density and admission rates for two low-variation conditions, the basic access index—which includes perceived access to healthcare and other basic needs—remained associated with lower rates of hospitalisation.

While the basic access index and the physical health domains of well-being are expected to be associated with fewer hospitalisations, other domains of the WBL, including better emotional health and overall life evaluation, were also associated with lower hospitalisation rates. The strength of the inverse association between emotional health and hospitalisation rates was similar to that of physical health and the composite WBL. These findings support the hypothesis that domains of well-being beyond physical health and access to basic needs should not be overlooked in their role supporting population health.

Our study has several limitations that should be considered when interpreting results. First, this was a cross-sectional study, so no implications about the directionality of the observed association or causation may be inferred. Though we reviewed the evidence supporting the hypothesis that higher community well-being leads to lower hospitalisation rates, it is also possible that frequent hospitalisations result in depressed well-being. However, hospital admissions affect only a portion of the community population, so it is less likely to be able to influence community well-being. It is also possible that an unknown confounding variable explains the observed inverse relationship between higher well-being and lower hospitalisation rates. However, we controlled for multiple healthcare access factors, as well as demographic variables, making it less likely. In addition, because we were limited to zip codes that had at least 10 participants in the WBI survey, we were only able to include more populated zip codes. As such, our findings may be less representative of rural or less densely populated communities. Further, we were unable to account for disease prevalence in our study, which may impact hospitalisation rates. However, we did control for hospitalisation for low-variation conditions, and the states included do not have known marked differences in the prevalence of cardiovascular or respiratory diseases. Finally, though we tried to include a diverse sample of states, we were limited to those that had statewide all-payer claims data available and that we had access to. We did not include states from the Southeast nor the West Coast, so findings may be less generalisable to those regions.

CONCLUSION

Higher community well-being was associated with lower rates of all-cause, cardiovascular and respiratory hospital admissions, even after adjusting for sociodemographic and healthcare intensity factors. Though it remains to be tested, these results suggest that policies that promote community well-being may also reduce hospitalisation rates.

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Data availability statement Data are available in a public, open access repository. A de-identified data set with zip-code resident well-being data from Gallup-Sharecare linked to all-cause hospitalization rates is available on ICSPR Open at: https://www.openicpsr.org/openicpsr/project/115103/version/V1/view.

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