Refining the clinical application of the consideration of future consequences scale -14

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Refining the clinical application of the consideration of future consequences scale -14

Andrew Percy ¹ · Michael T. McKay ² · John L. Perry ³ · Jon C. Cole ²

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Abstract
Within the temporal psychology literature, evidence has been presented for a relationship between a variety of health indicators, including alcohol-related problems, as well as symptoms of anxiety and depression, and a variety of temporal constructs. Recently, it was claimed that the two factors of the Consideration of Future Consequences (CFC) Scale-14 (CFCS-14) were both practically and conceptually useful in understanding the relationship between CFC and various criterion variables. The present study examined the relationship between these two factors (consideration of immediate and consideration of future consequences) and symptoms of anxiety and depression, as well as problematic alcohol use. Participants were recruited from a University in the North West of England, and completed the CFCS-14, the Alcohol Use Disorders Identification Test, and the Hospital Anxiety and Depression Scale. Results offer support for the psychometric validity and internal consistency of the CFCS-14, and further reveal a modest relationship between its factors and criterion variables. Compared to temporal psychology measures focusing on time attitudes or time perspective, the clinical or practical utility of the CFCS-14 in understanding these health domains may be limited.

Keywords Consideration of future consequences · Anxiety · Depression · Alcohol · AUDIT · Time perspective

Introduction
Given that symptoms of psychological distress, including anxiety and depression have been observed in university undergraduates (e.g., Blanco et al. 2008; Ibrahim et al. 2013; Vazquez and Blanco 2008), it seems important to investigate both the correlates of these symptoms, and potential means by which these might be ameliorated. In this context the area broadly known as ‘temporal psychology’ is potentially useful as it has been theorized that (specifically) future-oriented constructs, such as hope and optimism, are developmentally adaptive (e.g., Burrow et al. 2010; Schmid and Lopez 2011; Sun and Shek 2013), and are associated with lower levels of psychiatric symptomatology.

Temporal psychology in its broadest sense examines the extent to which thoughts about, and feelings towards the past, present and future, influence human behavior. However, temporal psychology is a multi-faceted area of research within which narrower or more nuanced dimensions have been widely studied. These include, but are not limited to, time perspective, time attitudes, temporal focus, and consideration of future consequences (CFC).

Time perspective, as conceptualized in the Zimbardo Time Perspective Inventory (ZTPI), is said to be “the often nonconscious process whereby the continual flows of personal and social experiences are assigned to temporal categories, or time frames, that help to give order, coherence, and meaning to those events” (Zimbardo and Boyd 1999, p.1271). The ZTPI assesses time perspective in five domains, past positive, past negative, present hedonistic, present fatalistic, and future.
problems (Chavarria et al. 2015; Lindene et al. 2014). In addition, others have shown that a past negative sequence and a lower future time perspective and more problematic alcohol use (e.g., Beenstock et al. 2011; McKay et al. 2018) called for greater clarity and specificity in the description of temporal constructs, and rather than using macro-level terms such as ‘future orientation’, called on researchers to be more precise in their description of scales and constructs.

Within the broader temporal psychology literature there is reason to believe that there is a meaningful relationship between alcohol use, psychopathology and temporal constructs. A number of studies have reported a significant and meaningful relationship between lower consideration of future consequences and a lower future time perspective and more problematic alcohol use (e.g., Beenstock et al. 2011; McKay et al. 2014). In addition, others have shown that a past negative orientation can be associated with greater alcohol-related problems (Chavarria et al. 2015; Linden et al. 2014).

A number of studies have reported meaningful relationships between mental health indicators and various dimensions of temporal psychology. For example, Van Beek et al. (2011), using the ZTPI reported that a past positive (negatively), past negative and present fatalistic (both positively) temporal bias were indicative of psychiatric problems, including depression and suicidal ideations. Elsewhere, past negative has been shown to be positively correlated with depression (Zimbardo and Boyd 2008; Lyubomirsky and Nolen-Hoeksema 1995). Elevated levels of present fatalistic time perspective were also reported as being significantly associated with psychopathology, including heightened risk of suicidal ideation, less extraversion and consciousness, and more neuroticism (Laghi et al. 2009; Van Beek et al. 2011).

The Consideration of Future Consequences Scale (CFCS; Strathman et al. 1994) was developed to assess individual differences in the degree to which people consider the potential future outcomes of behavior, and the degree to which that consideration affects present behavior. The 12-item CFCS was originally developed as a unidimensional scale, so that those scoring high on the CFCS-12 were considered to be high in CFC, with the opposite true for those scoring low (for a review see Joireman and King 2016). However, a range of studies have reported a variety of optimal factor solutions for the scale (Hevey et al. 2010; Joireman et al. 2008; McKay et al. 2015a, b; Strathman et al. 1994).

In an attempt to create a more psychometrically valid two-factor solution, Joireman et al. (2012) developed the CFCS-14 by adding two additional future oriented items. Hence, the CFCS-14 was purposively developed to simultaneously assess CFC-Future (CFC-F; consideration of future consequences) and CFC-Immediate (CFC-I; consideration of present consequences, where items are no longer reverse scored). The developers argued that as well as providing a more psychometrically reliable and internally consistent two-factor scale, a two-factor solution would be conceptually and practically useful in understanding the relationship between consideration of future consequences and criterion variables (Joireman et al. 2012). Indeed, McKay et al. (2016b) reported a good fitting two-factor solution for the CFCS-14, but these authors questioned the practical utility of a two factor solution specifically in relation to alcohol-use and scores on symptoms of psychopathology, claiming that the CFC-F factor was largely redundant in explaining any variance in these scores.

Although the study of temporal psychology is not new (e.g., Lewin 1942), an increasing body of evidence is beginning to demonstrate the potential utility of the construct, in particular of future orientation, in both general health (e.g., Hall et al. 2014) and substance misuse treatment (e.g., Davies and Filippopoulos 2015). This is potentially significant given the fact that the inability to consider the future consequences of behavior has been widely linked to health-compromising behaviors (e.g., Apostolidis et al. 2006; Daugherty and Brase 2010; Fieulaine and Martinez 2011). A scale that can simultaneously assess consideration of future consequences (CFC-F) and consideration of immediate consequences (CFC-I) would seem to be of considerable value to those planning health promotion and/or therapeutic interventions. To date, with only one study having investigated the relationship between the CFCS-14 scores and self-reported psychopathological symptoms (McKay et al. 2016b), this conceptual utility argument cannot be substantiated for these constructs.

Despite the promise offered by the CFCS-14 in terms of psychometric validity, internal consistency and conceptual utility, one issue of concern remains. A number of studies have reported the CFC-F factor to be unrelated to criterion variables including environmental concern, alcohol-related problems, and psychopathology (Arnocky et al. 2014; McKay et al. 2016a). Additionally, Van Beek et al. (2013), using a domain-specific CFC scale showed that CFC-I food predicted eating behaviors, while CFC-F exercise predicted exercising behaviors. The totality of these results point to a complex, but
evolving CFC literature. Despite the fact that unlike the ZTPI, the AATI-TA, or the TAS, the CFCS-14 lacks a past element, results of studies described above illustrating positive outcomes for more future-focused individuals, it might reasonably be lower levels of anxiety and depressive symptomatology would be associated with greater consideration of future consequences.

In the only study to date to examine the relationship between CFCS-14 scores and symptoms of anxiety and depression, McKay et al. (2016b) reported no significant relationship between CFC-I ($p = .462$) nor CFC-F ($p = .115$) scores and symptoms of depression, and a significant ($p = .033$) relationship between CFC-I and symptoms of anxiety, but no significant relationship between CFC-F ($p = .461$) and anxiety. However, it should be pointed out that the standardized beta coefficient for these findings did not reach Ferguson’s minimum practical effect size ($\beta \geq .20$; Ferguson 2009). Despite the fact that McKay et al. (2016b) reported a good fitting and internally consistent two-factor model, their results suggested that, in terms of CFC-F and CFC-I, there was only limited conceptual utility in terms of understanding these particular health-related criterion variables.

Given the suggestion that the CFCS-14 is psychometrically more robust than the CFCS-12, and the suggestion that it has both conceptual and practical utility, the present study had two specific aims. Firstly, we aimed to undertake a further examination of the factor structure of the CFCS-14, and secondly, to re-examine the clinical utility of the best fitting model with regard to problematic alcohol use, as well as symptoms of anxiety and depression. This is important in a literature replete with psychometric and conceptual concerns where replication of studies is rare, and this study is only the second to examine how CFCS-14 scores relate to these criterion variables.

**Method**

**Participants**

Participants were 369 adults (aged 18–40 [mean (+SD) 21.68 (4.17)]; 46.3% male), recruited in a University in the North West of England. Participants completed all measures using pen and paper format. No incentives were offered for participation and completion took between 25 and 30 min. The study was given ethical approval by the relevant university ethics committee and all participants gave informed consent.

**Measures**

The CFCS-14 (Joireman et al. 2012) is made up of seven CFC-F items (e.g., *When I make a decision, I think about how it might affect me in the future*), measuring consideration of future consequences, and seven CFC-I items (e.g., *I only act to satisfy immediate concerns, figuring the future will take care of itself*), measuring consideration of immediate consequences. Responses were on a 7-point Likert-type scale from 1 (*very unlike me*) to 7 (*very like me*). In their development of the scale Joireman et al. (2012) reported two highly reliable ($\alpha$) factors; CFC-Future = .80 and CFC-Immediate = .84, as well as a small-sized (Ferguson 2009) correlation coefficient between the two factors ($r = -.27$).

The Hospital Anxiety and Depression Scale (HADS; Zigmond and Snaith 1983) yields scores for anxiety (HADS-A) and depression (HADS-D) on separate subscales with scores ranging from zero to twenty-eight, with a higher score indicating a greater degree of anxiety or depression.

The Alcohol Use Disorders Identification Test (AUDIT; Saunders et al. 1993) is a 10-item questionnaire with valid and reliable scores across different contexts and cultures (e.g., De Meneses-Gaya et al. 2009). When used to detect problematic alcohol use in a population of university undergraduates, AUDIT demonstrated good sensitivity (.94) and specificity (.92; Adewuya 2005). In the present study scores on all 10 items were summed to produce an overall AUDIT score.

**Statistical Analyses**

Three models were estimated using the maximum likelihood estimator with standard errors and chi-square test statistics robust to non-normality (MLR) in Mplus (Muthén and Muthén 2012). First, we estimated a unidimensional model in which all 14 items loaded onto a single factor. Second, we estimated a two-factor model in which the factor correlation was constrained to zero. Items 1, 2, 6, 7, 8, 13 and 14 were assigned to the first factor (i.e., CFC-F), and Items 3, 4, 5, 9, 10, 11, and 12 were assigned to the second factor (i.e., CFC-I). Third, we estimated the same two-factor model described earlier with the factor correlation freely estimated. For each estimated model, we examined the following fit indexes: the comparative fit index (CFI), the Tucker–Lewis index (TLI), the root mean square error of approximation (RMSEA) and its 95% confidence interval, and the root mean square residual (SRMR).

Values of CFI and TLI at or above .95 and values of RMSEA at or below .05 indicate close model fit. CFI and TLI values between .90 and .95 and RMSEA values between .05 and .08 indicate acceptable model fit. Values of SRMR at or below .08 indicate close model fit (Hu and Bentler 1999). Regression models were then estimated, in which a range of outcomes (AUDIT, HADS-A, HADS-D scores) were regressed onto the latent factors and selected covariates (age and gender). Further, because the study was exploratory, we also used Binary Logistic models to assess the
relationship between Anxiety and Depressive caseness (HADS >8; Snaith and Zigmond 1994), problematic alcohol use (AUDIT >8; Beenstock et al. 2011) and CFC-I and CFC-F scores.

To aid the interpretation of the regression analyses, we applied the criteria of Ferguson (2009). Accordingly, a recommended minimum effect size for standardized $\beta$ (or a practically significant effect size) is $\geq .2$, a moderate effect size $\geq .5$, and a strong effect $\geq .8$.

### Results

Table 1 displays the model fit results for a variety of solutions. Results clearly demonstrate that the fit indices for both the unidimensional model and the two uncorrelated factors model were suboptimal. The fit indices for the two factor solution were acceptable. Table 2 displays the factor loadings for the two-factor model. It is observed that both CFC-I and CFC-F items load strongly on their respective factors.

### Table 1

| Model Type          | $\chi^2$ | df  | CFI  | TLI  | RMSEA | RMSEA 90% CI | SRMR |
|---------------------|----------|-----|------|------|-------|---------------|------|
| Unidimensional      | 646.060*** | 77  | 0.590 | 0.516 | 0.142 | 0.132–0.152   | 0.117 |
| Uncorrelated 2 factor | 231.718*** | 77  | 0.889 | 0.868 | 0.074 | 0.063–0.085   | 0.129 |
| Correlated 2 factor  | 183.281*** | 76  | 0.923 | 0.908 | 0.062 | 0.050–0.073   | 0.053 |

CFI, Comparative Fit Index; TLI, Tucker Lewis Index; RMSEA, Root Mean Square Error of Approximation; SRMR, Standardized Root Mean Square Residual; CI, Confidence Interval; *** $p < .001$

### Table 2

| Item                                                                 | CFC-F B | CFC-F SE | CFC-I B | CFC-I SE |
|---------------------------------------------------------------------|--------|---------|--------|---------|
| CFC-F 1. I consider how things might be in the future, and try to influence those things with my day to day behavior. | 0.851  | 0.058   |        |         |
| CFC-F 2. Often I engage in a particular behavior in order to achieve outcomes that may not result for many years. | 0.969  | 0.069   |        |         |
| CFC-F 6. I am willing to sacrifice my immediate happiness or wellbeing in order to achieve future outcomes. | 0.876  | 0.068   |        |         |
| CFC-F 7. I think it is important to take warnings about negative outcomes seriously even if the negative outcome will not occur for many years. | 0.621  | 0.078   |        |         |
| CFC-F 8. I think it is more important to perform a behavior with important distant consequences than a behavior with less important immediate consequences. | 0.590  | 0.060   |        |         |
| CFC-F 13. When I make a decision, I think about how it might affect me in the future. | 0.791  | 0.057   |        |         |
| CFC-F 14. My behavior is generally influenced by future consequences. | 0.871  | 0.064   |        |         |
| CFC-I 3. I only act to satisfy immediate concerns, figuring the future will take care of itself. |        |         | 1.175  | 0.053   |
| CFC-I 4. My behavior is only influenced by the immediate (i.e., a matter of days or weeks) outcomes of my actions. |        |         | 1.036  | 0.064   |
| CFC-I 5. My convenience is a big factor in the decisions I make or the actions I take. |        |         | 0.631  | 0.081   |
| CFC-I 9. I generally ignore warnings about possible future problems because I think the problems will be resolved before they reach crisis level. |        |         | 0.923  | 0.068   |
| CFC-I 10. I think that sacrificing now is usually unnecessary since future outcomes can be dealt with at a later time. |        |         | 0.860  | 0.070   |
| CFC-I 11. I only act to satisfy immediate concerns, figuring that I will take care of future problems that may occur at a later date. |        |         | 1.114  | 0.052   |
| CFC-I 12. Since my day to day work has specific outcomes, it is more important to me than behavior that has distant outcomes. |        |         | 0.481  | 0.068   |

$B$, unstandardized coefficient; $SE$, standard error of $B$; All estimates are statistically significant ($p < .001$) except *$p < .01$ and † non-significant. CFC-I, Consideration of Immediate Consequences; CFC-F, Consideration of Future Consequences. Correlation between CFC-I and CFC-F, $r = −.37$
Internal consistency (alpha) estimates were all satisfactory in the present study and were as follows: (CFC-F, $\alpha = .82$; CFC-I, $\alpha = .82$; HADS-A, $\alpha = .81$; HADS-D, $\alpha = .70$; AUDIT, $\alpha = .83$). Table 3 displays the regression coefficients of individual regression models for AUDIT, HADS-A and HADS-D scores on the two correlated factors (CFC-F and CFC-I). Additional covariates included age and gender. Both CFC-F and CFC-I statistically predicted AUDIT scores with higher CFC-F and lower CFC-I significantly associated with higher AUDIT score. Both factors also predicted HADS-A, but this time both higher CFC-F and higher CFC-I predicted higher anxiety scores. For HADS-D, CFC-F was non-significant, while CFC-I was a significant predictor, with higher CFC-I predicting higher depression scores.

Finally, Table 4 displays the results of three Binary Logistic models assessing the relationship between CFC-I and CFC-F scores and caseness for all dependent measures, adjusted for age and gender. Results show a similar pattern to those in Table 3, so that the $R^2$ values suggest that CFC generally explains a modest amount of variance in caseness for problematic alcohol use, anxiety, and depression.

### Table 3  Regression parameter estimates (two correlated factors; CFC-F, CFC-I)

|              | $\beta$ | SE  | p    |
|--------------|---------|-----|------|
| **AUDIT, $R^2 = 0.202$**         |         |     |      |
| Intercept    | 2.904   | 1.665 | <0.001 |
| Age          | 0.171   | 0.074 | <0.001 |
| Gender       | 0.479   | 0.368 | <0.001 |
| CFC-F        | 0.222   | 0.405 | <0.001 |
| CFC-I        | 0.122   | 0.391 | 0.041 |
| **HADS-A, $R^2 = 0.049$**        |         |     |      |
| Intercept    | 2.134   | 1.135 | <0.001 |
| Age          | 0.112   | 0.051 | 0.032 |
| Gender       | 0.222   | 0.423 | 0.032 |
| CFC-F        | 0.143   | 0.252 | 0.020 |
| CFC-I        | 0.154   | 0.241 | 0.009 |
| **HADS-D, $R^2 = 0.027$**        |         |     |      |
| Intercept    | 0.894   | 0.901 | <0.001 |
| Age          | 0.058   | 0.040 | 0.264 |
| Gender       | 0.036   | 0.335 | 0.731 |
| CFC-F        | 0.098   | 0.204 | 0.124 |
| CFC-I        | 0.170   | 0.193 | 0.005 |

$\beta$, Standardized coefficient; $B$, Unstandardized coefficient; SE, Standard error of B; For continuous covariates STDYX method of standardisation was used. For binary covariates STDY was used. AUDIT, Alcohol Use Disorders Identification Test; HADS, Hospital Anxiety and Depression Scale (A, Anxiety; D, Depression)

### Table 4  Summary of logistic regression analysis of the relationship between CFC scores and caseness for Alcohol-related problems, Anxiety, and Depression

|              | OR (95% CI) | Wald $\chi^2$ | p-value | Nagelkerke $R^2$ |
|--------------|-------------|---------------|---------|------------------|
| **AUDIT**    |             |               |         |                  |
| Gender       | 2.29 (1.40, 3.73) | 11.02 | 0.001 |
| Age          | 0.91 (0.87, 0.96) | 11.01 | 0.001 |
| CFC-I        | 1.58 (1.20, 2.06) | 10.86 | 0.001 |
| CFC-F        | 0.73 (0.54, 0.99) | 4.08  | 0.043 |
| Constant     | 9.91        | 3.92          | 0.048  |
| **Anxiety**  |             |               |         |                  |
| Gender       | 0.64 (0.41, 0.98) | 4.21  | 0.040 |
| Age          | 0.94 (0.89, 0.99) | 4.25  | 0.039 |
| CFC-I        | 1.32 (1.04, 1.68) | 5.20  | 0.023 |
| CFC-F        | 1.44 (1.10, 1.90) | 6.88  | 0.009 |
| Constant     | 0.19        | 2.24          | 0.134  |
| **Depression**|            |               |         |                  |
| Gender       | 0.86 (0.49, 1.51) | 0.29  | 0.590 |
| Age          | 1.05 (0.98, 1.11) | 2.14  | 0.144 |
| CFC-I        | 1.10 (0.81, 1.50) | 0.37  | 0.541 |
| CFC-F        | 1.23 (0.86, 1.75) | 1.31  | 0.252 |
| Constant     | 0.02        | 7.97          | 0.005  |

CFC, Consideration of Future Consequences (F, Future; I, Immediate); OR, Odds Ratio; CI, Confidence Interval

### Discussion

The present study examined the psychometric validity, internal consistency and clinical utility in terms of psychiatric symptomatology, of scores on the CFCS-14 using a university-based sample in the UK. Overall, results show that the model fit for the two-factor model was acceptable. There was no support for a unidimensional CFCS-14 scale, and accordingly, the work undertaken by Joireman et al. (2012) to develop a valid two-factor scale is supported.

In terms of the conceptual or clinical utility of the CFCS-14 in understanding symptoms of anxiety and depression, as well as problematic alcohol use, results revealed a number of issues. Firstly, the r-square values in all models were quite low, both for raw scale scores, and when AUDIT and HADS scores were categorized in terms of clinical caseness. This suggests that the variance in these scores explained by CFCS-14 scores is relatively low overall. Ferguson (2009) suggested that only $R^2$ values of $\geq 0.04$ could be interpreted as being meaningful, and accordingly, the model for HADS-D scores was sub-optimal. Additionally, effect sizes for all but one of the significant results did not reach Ferguson’s (2009) threshold ($\beta \geq 0.2$) for practical significance. In other words, while the $p$ values may have suggested a significant relationship, the effect sizes question how meaningful that relationship actually is. The only CFCS-14 standardized beta value that reached practical
significance was the CFC-F value in the AUDIT model (β = −.22), and this result supports a wider literature suggesting that higher levels of alcohol use and/or alcohol-related problems are related to a lack of future focus, orientation or planning (Cole et al. 2016; McKay et al. 2014; Zimbardo and Boyd 2008).

Where, in other studies the use of scales measuring time attitudes (the affective dimension) and time perspective (a mixture of affective, cognitive and behavioral items) have yielded meaningful results for symptoms of anxiety and depression, the CFCS-14 (consisting of cognitive and behavioral items) does not. Indeed, closer examination of the CFCS-14 items (Table 2) reveals that they lack both valence (positive or negative) and specificity. This is in contrast to both the ZTPI (Zimbardo and Boyd 1999), and the Adolescent Time Inventory-Time Attitudes Scale (ATI-TA; Mello and Worrell 2007) where items are either framed negatively or positively, or are specific to particular behaviors. In addition, both the ZTPI and the ATI-TA have more subscales (five and six respectively) and are therefore potentially more sensitive to variation in criterion variable scores. It is perhaps for this reason that the relationship between psychiatric symptoms, alcohol use, and temporal psychology emerged more obviously with the use of these scales (ZTPI, McKay et al. 2016a; ATI-TA, McKay et al. 2017) than with the CFCS-14. This is not a criticism of the CFCS-14 scale per se, more an observation about the CFC construct. Where previously McKay et al. (2016b) concluded that the CFCS-14 was limited in its conceptual utility in respect of HADS-A and HADS-D scores, a better interpretation of those results, and the results of the present study combined, might be that consideration of the future as assessed by the CFCS-14 is practically unrelated to the symptoms of anxiety and depression, and poorly related to alcohol-related problems, in young people. Further, as this developing literature begins to grow in terms of numbers of studies, a review of studies using a range of temporal psychology instruments, and in particular the effect sizes reported therein, may be more instructive in terms of what measures relate best to alcohol-related measures, as well as measures of psychopathology.

Secondly, insofar as these modest results are practically applicable, they point to the fact that higher levels of depressive symptomatology are significantly associated with consideration of the immediate future (essentially present orientation), but not to consideration of the more distant future. Conversely, the results point to the fact that higher levels of anxiety symptomatology are significantly related to higher levels of consideration of both the immediate and more extended future. This is somewhat different to the results previously reported by McKay et al. (2016b), and more in keeping with a conceptual utility argument (the fact that both CFC-I and CFC-F are significantly associated with outcomes). However, the small beta values limit the practical significance of any findings (Ferguson 2009).

One practical implication of these findings relates to the potential utility of the CFCS-14 in clinical settings. Using this scale, it is questionable, given the overall amount of variance explained, whether the CFCS-14 might help with the understanding of symptoms of anxiety and depression, or problematic alcohol use. This seems particularly important in view of recent developments where temporal psychology is being employed in clinical settings (Davies and Filippopoulos 2015). However, this is only the second study to examine this relationship, and both this and the previous study (McKay et al. 2016b) were cross sectional, and used a University sample. Indeed, further studies in a range of population types are required to be able to draw more definitive conclusions.

The present study is not without limitations. Firstly, all data were self-reported. Secondly, participants were a university sample, and therefore, it is not clear if results will generalise to the general population. Finally, this is only the second study to examine associations between the CFCS-14 and symptoms of psychopathology. Accordingly, more work will need to be undertaken in the examination of these associations in diverse samples in different cultural contexts.

In conclusion, the present study adds to the growing literature supporting the psychometric validity and internal consistency of the two-factor CFCS-14, and that it should be scored and applied accordingly. However, the amount of variance explained in this context is quite modest, and researchers examining the relationship between temporal psychology and symptoms of anxiety and depression, or alcohol use, might be better advised to focus attention on the affective dimension (time attitudes).

Compliance with Ethical Standards

Conflict of Interest On behalf of all authors, the corresponding author states that there is no conflict of interest.

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