In this paper, we examine the Other as Shamer Scale-2 (OAS-2), a unidimensional and brief scale to directly assess external shame. In three studies with three independent samples of a Turkish university, we present evidence for OAS-2 validity with respect to well-being outcomes (subjective happiness, flourishing, and subjective well-being) and psychological distress outcomes (depression, anxiety, stress, and loneliness) through direct comparisons with existing measures.

In Study 1 (N = 311), confirmatory factor analyses, measurement invariance across gender, and Item Response Theory (IRT) were examined. In Study 2 (N = 380), criterion-related validities of the OAS-2 were analyzed. In Study 3 (N = 252), incremental validity was examined using PROCESS. Also, internal consistency, composite reliability, and temporal reliability (n = 89) of the OAS-2 were investigated.

In Study 1, confirmatory factor analyses supported the unidimensionality of the measure. The results provide support for measurement invariance across gender. All item scores fit the IRT model and were fit with ordered, progressing hierarchies in their step difficulties. In Study 2, criterion-related validity for the OAS-2 was demonstrated through positive correlations with loneliness, and negative correlations with subjective happiness and flourishing. In Study 3, findings indicated the mediation impact of external shame on well-being via psychological distress. The OAS-2 showed satisfactory reliability coefficients.

Overall, the OAS-2 proved to be a valuable and reliable tool, which presents a short form to measure external shame. In addition, it was observed that the OAS-2 was related to both well-being and psychological distress.

**KEY WORDS**
external shame; well-being; psychological distress; measurement invariance; item response theory
BACKGROUND

Shame is a powerful, self-conscious emotion that emerges from the experience of being seen as defective, inadequate or powerless by others. Gilbert’s model “The Evolutionary and Biopsychosocial Model of Shame” provides significant information about the formation and development of shame (Gilbert, 2009, 2010). According to this model, from the moment they are born, individuals need to create positive emotions about themselves in the minds of others. Therefore, the way individuals experience interpersonal relationships at an early age has a very important effect on how we think we exist in the minds of others. Accordingly, shame emerges from our complex, evolved mental abilities to be aware of “how we exist for others” and to make predictions regarding what they think and feel about us (Gilbert, 2002, 2003). Also, shame has been described as both an intrapsychic and an interpersonal emotion (Tangney, 1995). On the other hand, shame conveys a strong message of self-abasement (Giner-Sorolla, Kamau, & Castano, 2010). In addition, shame, which develops as a system which regulates psychobiological reactions, may lead to rejection, ostracism, put-downs and even persecution (Saggio et al., 2017). From this point of view, the concern of being seen as inferior to others in an unattractive way can be experienced (Gilbert, 2007). For this reason, shame can be accepted as an indicator or stimulus of potential social damage.

Shame can be divided in two categories: internal and external (Gilbert, 1998, 2003; Gilbert & Procter, 2006). Internal shame is defined as a negative self-evaluation focusing on personal mistakes and perceived deficits (Gilbert, 1998). The individual maintains severe self-criticism processes (Gilbert, 2007). In this context, internal shame is the tendency to see oneself as a worthless person (Marta-Simões, Ferreira, & Mendes, 2016). On the other hand, external shame is the shame we experience when we believe we are held negatively in the minds of others (Matos, Pinto-Gouveia, & Costa, 2013). According to another definition, external shame includes the negative opinions of the self as seen from the perspective of others (Kim, Thibodeau, & Jorgensen, 2011). In external shame, the attentions of individuals are focused on the minds of others and their behaviors are orientated toward trying to influence their images in the minds of others (Matos et al., 2013). In other words, in external shame, the attention is focused on the minds of others with an aim to imagine what is in their minds and thus exhibit behaviors meant to change their opinions (Balsamo et al., 2015).

EXTERNAL SHAME, LONELINESS, WELL-BEING, AND PSYCHOLOGICAL DISTRESS

For shame, the term “sleeper of psychopathology” is used (Lewis, 1987, p. 1). Not only are there efforts to examine what may elicit shame, but there is also a growing interest in investigating the association between shame and individuals’ well-being, including psychological distress.

Because shame is a self-disapproving, devastating and painful experience, it makes individuals vulnerable to psychological distress (Lee, Anderson, & Klimes-Dougan, 2016). Moreover, shame includes subjective emotions which make individuals feel small, insignificant, helpless and worthless, which can cause them to experience distress (Kim et al., 2011). Therefore, shamed individuals can be passive in the face of psychological distress such as depression, anxiety, and stress. In their longitudinal research, Cunha and her colleagues (2016) found that external shame strongly predicted higher levels of depression. Also, it was specified that shame was a factor in increasing depression, according to the modelling (Lee et al., 2016; Mereish & Poteat, 2015; Thoresen, Aakvaag, Strom, Wenzel-Larsen, & Birkeland, 2018). In a meta-analysis study, external shame was associated with increasing depression (Kim et al., 2011). In addition, it was reported in other studies that shamed individuals experienced more anxiety, and stress (Pineles, Street, & Koenen, 2006; Mensinger, Tylka, & Calamari, 2018; Thomson & Jaque, 2018). Lastly, Castilho and her colleagues (2017) determined that external shame could directly increase depression, anxiety, and stress. Therefore, we hypothesized that external shame may have a positive relation with psychological distress (i.e. depression, anxiety, and stress).

There is expected to be a negative relationship between well-being and shame. Shame can reduce well-being because it is an unpleasant and painful emotion (Cibich, 2016). Similarly, shame is a negative emotion which decreases well-being (Cibich, Woodyatt, & Wenzel, 2016). Experiences with shame can negatively affect self-engagement, social comparison processes, subsequent self-esteem, and emotional well-being. In studies, the negative relationship between shame and well-being is emphasized (e.g., Cavalera et al., 2017; Choma, Shove, Busseri, Sadava, & Hocker, 2009; Kelly & Carter, 2013). In addition, shame decreases well-being according to the model presented by Varghese (2015). In fact, shame can negatively affect the well-being of individuals by leaving them unprotected against forms of psychological distress such as depression, anxiety, etc. (Clapton, Williams, & Jones, 2018). Other researchers have also revealed that shame has a negative relationship with life satisfaction and happiness (e.g., Bugay & Demir, 2011; Sousa et al., 2019). Based on these explanations, we also hypothesized that there would be a negative relationship between shame and well-being.

Because external shame may lead to disengagement and self-effacement, it can cause individuals to feel lonely and isolated (Thoresen et al., 2018). In addition, because shamed individuals can be anxious about social assessment, potential rejections or regressions,
they can withdraw from social interactions because of confidence issues (Breggin, 2015; Clapton et al., 2018; Kaufman, 2004). Furthermore, due to shame’s potential for preventing the establishment of close relationships, individuals can have higher levels of loneliness (Teroni & Deonna, 2008). Similarly, shamed individuals can feel lonely because they experience feelings of worthlessness about being engaged with others (Jordan, 2004). In addition, conceptual and empirical works emphasize that shame can increase loneliness (e.g., Mereish & Poteat, 2015; Thoresen et al., 2018; Rostami & Jowkar, 2016). Therefore, we hypothesized that external shame and loneliness could be positively related to each other based on both theoretical and empirical research results.

MEASUREMENT OF EXTERNAL SHAME

The Other as Shamer Scale (OAS) is commonly used to measure external shame. The OAS was developed by Allan, Gilbert, and Goss (1994) and Goss, Gilbert, and Allan (1994) based on the Internal Shame Scale of Cook (1994). In the first OAS to be developed, there are 18 items, and the items contain statements starting with “Other people see me as...”. Validity studies were conducted with university students, and a three-factor structure was revealed (Goss et al., 1994). Although the OAS is three-dimensional (inferiority, emptiness, and mistakes), it is understood that studies are frequently conducted through an examination of the total external shame score (e.g., Gilbert & Miles, 2000; Matos et al., 2013; Matos, Pinto-Gouveia, Gilbert, Duarte, & Figueiredo, 2015).

It can be stated that the validity and the reliability of the OAS have been tested in different populations up to the present. For example, Balsamo and colleagues (2015) confirmed the structure of the OAS in an Italian sample. Similarly, the OAS has been adapted and used in the British (Wood & Irons, 2017), Portuguese (Matos, Pinto-Gouveia, & Duarte, 2012), Greek (Kotrotsiou et al., 2017), and Scottish (Laithwaite et al., 2009) cultures.

Recently, an abbreviated form of the OAS (OAS-2) was developed (Matos et al., 2015). The OAS-2 consists of eight items selected from the OAS. As a result of the OAS-2’s confirmatory factor analysis, its one-dimensional structure was approved. It was found that acceptable internal consistency had a significant relationship with internal shame, psychopathology, and anger. Therefore, a valid, reliable and economic measurement tool for evaluating external shame emerged (Matos et al., 2015). The psychometric strength of the OAS-2 was proven in Italian university students (Saggino et al., 2017), Portuguese adolescents (Cunha, Xavier, Cherpe, & Gouveia, 2017) and Portuguese children (Benevides, da Motta, Sousa, Caldeira, & Carvalho, 2016).

PURPOSE OF THE CURRENT STUDY

Although shame is a universal experience (Martasimoes et al., 2016), the concept can differ according to cultures (Gilbert, 1998, 2003). Although the Other as Shamer Scale-2 (OAS-2) is consistently used by researchers, the studies regarding its adaptation to other populations and its psychometric characteristics are not sufficient.

Most of the studies examining the validity and reliability of the OAS-2 were carried out in a western-individualist culture (Balsamo et al., 2015; Kotrotsiou et al., 2017; Laithwaite et al., 2009; Matos et al., 2015; Saggino et al., 2017). Therefore, in a collectivist culture, a measurement tool that can assess external shame is lacking. Hence, the study of external shame will contribute to a measurement tool that can be used in intercultural studies, and thus, to the conceptual development of shame.

Furthermore, although exploratory factor analysis (EFA) and confirmatory factor analysis (CFA), as well as concurrent, convergent and divergent validity, and internal and temporal reliability assessment, have been performed in the psychometric characteristics of the OAS-2 up to the present, its incremental, Item Response Theory (IRT) and composite reliability have not been carried out. Moreover, although measurement invariance across gender was conducted in the adolescent community (via independent sampling of referred and forensic samples), measurement invariance across gender was not tested in university students. Finally, the psychometric studies of the OAS-2 were conducted on a single participant group. Therefore, we aimed to carry out three different studies in this research. In Study 1, the OAS-2 was examined with CFA, measurement invariance, IRT and internal, composite and temporal reliability. In Study 2, the criterion-related validity of the OAS-2 was evaluated with loneliness and well-being. In Study 3, the incremental validity of the OAS-2 was evaluated with the model examining the mediating roles of depression, anxiety, and stress in the relationship between external shame and well-being.

STUDY 1

The original Other as Shamer Scale-2 from Matos and colleagues (2015) was translated into Turkish by three experts using the parallel blind technique. Then, back-translation was done according to Brislin’s (1980) suggestion. Subsequently, the translation forms were discussed with doctoral students and the final version of the Turkish OAS-2 was obtained.

In Study 1, the construct validity of the Turkish OAS-2 was examined with CFA and measurement invariance. In addition, Item Response Theory analyses were conducted to check item validity. After the
construct validity and IRT of the scale. Cronbach’s α, composite reliability, and test-retest were assessed in this study.

PARTICIPANTS

Three hundred eleven volunteer university students (159 females and 152 males) between ages 18 and 30 years ($M = 20.04$, $SD = 1.90$) were recruited. With regard to class levels, 27% were freshmen, 29.9% sophomore, 21.9% junior, and 21.2% senior. Table 1 shows the detailed participants’ demographics in all studies.

In Study 1, a new participant group was formed to examine the Turkish OAS-2’s test-retest reliability in order to investigate the degree to which the scale reflects a stable predisposition. This group consisted of 89 students. The age of the participants ranged from 18 to 24 years ($M = 19.96$, $SD = 1.81$).

PROCEDURE

The data were collected via the paper-pencil questionnaire format in the classroom environment. Written informed consent was obtained from the students before participation. In this respect, the questionnaire package was distributed only to volunteer students.

We performed CFA using AMOS Graphics to validate the factor structure of the Turkish OAS-2. As suggested in the literature (e.g., Hu & Bentler, 1999; Kline, 2015; Marsh, Hau, & Wen, 2004), we considered that goodness-of-fit index (GFI), normed fit index (NFI), and confirmatory fit index (CFI) ≥ .90, root mean square error of approximation (RMSEA) and standardized root-mean-square residual (SRMR) ≤ .08 were indicative of acceptable fit.

We also evaluated the factor structure for measurement invariance by gender to assess the equivalence of factor structure across female and male subjects. In this respect, configural, metric, scalar, and strict invariance analyses were tested. As suggested in the previous research literature (Chen, 2007; Cheung & Rensvold, 2002), $\Delta$CFI < .010 and $\Delta$RMSEA < .015 can be considered as indicating measurement invariance across different groups.

Lastly, we assessed the discrimination, difficulty, and informativeness of the scale with the Item Response Theory (IRT; Chalmers, 2012). In addition, IRT provides more detailed and reliable information at the individual and item level than Classical Test Theory. Within the IRT analysis, we used the item characteristic curve function and Graded Response Model, because of the polytomous nature (more than two answer categories) of the measure. IRT analyses were carried out with Stata 14.2.

RESULTS

We conducted a CFA using a maximum likelihood estimator to test the adequacy of the Turkish OAS-2 model and confirm the previously proposed Matos and colleagues’ OAS-2 model. Regarding the fit statistics, the chi-square statistic was significant, probably due to the size of the sample (Hair, Black, Babin, Anderson, & Tatham, 2014), but the ratio $\chi^2/df$ ($\chi^2/df = 2.81 \leq 5$), the SRMR (.039 ≤ .08), the RMSEA (.076 ≤ .08), the GFI (.955 ≥ .90), the NFI (.947 ≥ .90), and the CFI (.965 ≥ .90) were well inside the limits that allow the model to be accepted. All standardized factor loadings for the items were significant ($p < .001$), ranging from .54 (item 1) to .80 (item 6) (Table 2).

MEASUREMENT INVARINANCE

After the confirmation of the Turkish OAS-2, the scale was examined across gender. First, the CFA of the scale was made separately for female and male subjects. The results showed that the scale provided an adequate fit among both males [$\chi^2(20, n = 122) = 42.56$, $p < .05$; GFI = .935; NFI = .904; CFI = .946; SRMR = .052; RMSEA = .086] and females [$\chi^2(20, n = 159) = 48.21$, $p < .05$; GFI = .930; NFI = .926; CFI = .955; SRMR = .044; RMSEA = .084]. Then, tests of configural, metric, scalar, and strict invariance were performed. The results are presented in Table 3.
As can be seen in Table 3, the analysis of measurement invariance supported both configural and metric invariance (ΔCFI = .004; ΔRMSEA = .002) of the Turkish OAS-2 across gender, indicating that Turkish OAS-2 items have the same meaning for males and females. On the other hand, scalar (ΔCFI = .076; ΔRMSEA = .076) and strict invariance (ΔCFI = .012; ΔRMSEA = .004) models were not supported due to both goodness-of-fit and ΔCFI – ΔRMSEA values. In this regard, the use of scalar and strict invariance values should be cautious.

### Table 2
**CFA results of the Turkish OAS-2**

| Item | Standardized factor loading | Error variance | R² |
|------|----------------------------|----------------|----|
| 1    | .54                        | .71            | .43 |
| 2    | .69                        | .39            | .63 |
| 3    | .61                        | .61            | .65 |
| 4    | .62                        | .56            | .62 |
| 5    | .79                        | .26            | .38 |
| 6    | .80                        | .25            | .37 |
| 7    | .79                        | .30            | .48 |
| 8    | .65                        | .49            | .30 |

### Table 3
**Fit indices of gender invariance**

| Invariance          | χ²   | df  | GFI  | NFI  | CFI  | RMSEA | SRMR | ΔCFI | ΔRMSEA |
|---------------------|------|-----|------|------|------|-------|------|------|--------|
| Males               | 42.56| 20  | .935 | .904 | .946 | .08   | .05  | –    | –      |
| Females             | 48.20| 20  | .930 | .926 | .955 | .08   | .04  | –    | –      |
| Configural invariance | 89.83| 40  | .930 | .955 | .974 | .08   | .04  | –    | –      |
| Metric invariance   | 102.99| 47  | .923 | .949 | .971 | .08   | .06  | .004 | .002   |
| Scalar invariance   | 316.37| 62  | .776 | .867 | .895 | .16   | .21  | .076 | .076   |
| Strick invariance   | 376.42| 71  | .743 | .851 | .883 | .17   | .28  | .012 | .004   |

*Note. df – degrees of freedom; GFI – goodness-of-fit index; NFI – normed fit index; CFI – confirmatory fit index; RMSEA – root mean square error of approximation.*

### Table 4
**Item Response Theory parameter estimates for the Turkish OAS-2**

| Item | a   | b₁  | b₂  | b₃  | b₄  |
|------|-----|-----|-----|-----|-----|
| 1    | 1.59| −0.95| 0.41| 1.92| 3.27|
| 2    | 2.17| 0.02 | 1.21| 2.14| 3.63|
| 3    | 1.66| −0.44| 0.81| 2.07| 3.23|
| 4    | 1.85| 0.01 | 1.17| 2.13| 3.15|
| 5    | 2.96| 0.53 | 1.36| 2.16| 2.75|
| 6    | 3.37| 0.56 | 1.42| 2.06| 2.46|
| 7    | 2.76| 0.30 | 1.28| 1.89| 2.62|
| 8    | 1.90| 0.01 | 1.11| 2.23| 3.29|

**ITEM RESPONSE THEORY**

IRT, which is a popular method for evaluating educational tools, is currently being increasingly used in personality measures (e.g., Colledani, Anselmi, & Robusto, 2019; Coskun & Kara, 2019). These basic aspects of the IRT parameters can be visually shown in an S-shape curve known as the item characteristic curve (ICC). ICC analysis was performed with a Graded Response Model (GRM) because the Turkish OAS-2 has a five-point Likert type scale. Results of the IRT are presented in Figure 1 and Table 4.

As shown in Table 4, all a values are higher than 1.0. According to Baker (2001), an a value > 1.0 is considered highly discriminant. Therefore, IRT results indicated that items of the Turkish OAS-2 are capable of discriminating better performers from poor performers and possessing adequate item difficulty.
Figure 1. Item characteristics curve for the Turkish OAS-2.
RELIABILITY

Cronbach’s α, composite reliability (CR), and test-retest reliability were checked for the internal and the temporal reliabilities. The Turkish OAS-2 was reapplied to students with a 28-day interval for test-retest reliability. The result showed that the internal consistency reliabilities were highly acceptable (α = .88 and CR = .88). In addition, the test-retest analysis revealed that the Turkish OAS-2 showed temporal reliability with a correlation coefficient of .86.

STUDY 2

Study 2 was conducted with an aim to evaluate the criterion-related validity of the Turkish OAS-2. Data were collected on the Turkish OAS-2 along with three standardized validity measures whose details are given below in the measures section.

PARTICIPANTS

A total of 380 (202 female and 178 male) individuals participated in Study 2. They were aged 18-26 (average age: 20.45 years, SD = 2.05). With regard to class levels, 25.8% were freshmen, 28.2% sophomore, 23.7% junior, and 22.4% senior. Demographics are summarized in Table 1.

MEASURES

In addition to the Turkish OAS-2 (α = .89), the UCLA Loneliness Scale, the Subjective Happiness Scale, and the Flourishing Scale were added in this study. Information about the scales is presented below.

The UCLA Loneliness Scale (ULS-8) by Hays and DiMatteo (1987) comprises 8 items (e.g., “There is no one I can turn to”) in uni-dimensional structure. Participants responded on a four-point Likert scale ranging from 1 (I never feel this way) to 4 (I often feel this way). A sample item is: “In general, I consider myself.” The scores that can be obtained in the ULS-8 vary between 8 and 32. Rising scores indicate that the feeling of loneliness also increases. The ULS-8 was adapted to Turkish by Dogan, Akinci-Cotok, and Goçet-Tekin (2011) from the original version. The psychometric characteristics of the Turkish ULS-8 are satisfactory (α = .72; GFI = .97, AGFI = .94, CFI = .94, and RMSEA = .066; Dogan et al., 2011). In the current research, Cronbach’s α coefficient was .82.

The Subjective Happiness Scale (SHS) by Lyubomirsky and Lepper (1999) comprises 4-items in uni-dimensional structure. Participants responded on a four-point Likert scale ranging from 1 (I never feel this way) to 4 (I often feel this way). A sample item is: “In general, I consider myself.” The scores that can be obtained in the SHS vary between 4 and 28. Rising scores indicate that the level of subjective happiness also increases. The SHS was adapted to Turkish by Akin and Satici (2011) from the original version. The psychometric characteristics of the Turkish SHS are satisfactory (α = .70; NFI = .99, CFI = 1.00, RFI = .98, GFI = 1.00, AGFI = .99, and SRMR = .015; Akin & Satici, 2011). In the current research, Cronbach’s α coefficient was .73.

The Flourishing Scale (FS) by Diener and colleagues (2010) comprises 8-items in uni-dimensional structure. Participants responded on a seven-point Likert scale ranging from 1 (strongly disagree) to 7 (strongly agree). A sample item is: “I actively contribute to the happiness and well-being of others.” The scores that can be obtained in the FS vary between 8 and 56. Rising scores indicate a person with many psychological resources and strengths. The FS was adapted to Turkish by Telef (2013) from the original version. The psychometric characteristics of the Turkish FS are satisfactory (α = .80; GFI = .96, NFI = .94, RFI = .92, CFI = .95, IFI = .95, SRMR = .04; Telef, 2013). In the current research, Cronbach’s α coefficient was .88.

PROCEDURE

The same data collection process in Study 1 was also used in Study 2. Criterion-related criterion validity was assessed by computing Pearson’s correlations of the Turkish OAS-2 with the criterion measures as validity coefficients. Correlational analyses were conducted using IBM SPSS Statistics 22.

RESULTS

The results of the correlational analyses are presented in Table 5. The analyses showed significant positive correlations between the Turkish OAS-2 and loneliness (r = .48, p < .001). On the other hand, the Turkish OAS-2 is significantly and negatively correlated with subjective happiness, r = -.36, p < .001. A significant negative correlation was also observed between the Turkish OAS-2 and flourishing, r = -.39, p < .001.
STUDY 3

The incremental validity of the Turkish OAS-2 was also tested in a process model linking external shame to subjective well-being via depression, anxiety, and stress. Prior studies indicated that external shame may increase psychological distress (e.g., Castilho et al., 2017; Kim et al., 2011), and also may reduce well-being (Cibich et al., 2016; Varghese, 2015). Moreover, relevant literature showed that shame may reduce well-being via psychological distress (e.g., Clapton et al., 2018). Therefore, we tested the mediational model with external shame as an independent variable, depression, anxiety, and stress as mediator variables and subjective well-being as a dependent variable.

PARTICIPANTS

A total of 352 (181 female and 171 male) individuals participated in Study 3. They were aged 18-27 (average age: 20.14 years, SD = 1.98). With regard to class levels, 27% were freshmen, 25.6% sophomore, 22.4% junior, and 25% senior. Demographics are summarized in Table 1.

MEASURES

In addition to the Turkish OAS-2 (α = .89), the Depression, Anxiety, and Stress Scale, the Positive and Negative Affect Schedule, and the Satisfaction with Life Scale were used in this study. Information about the scales is presented below.

The Depression Anxiety Stress Scale (DASS) by Lovibond and Lovibond (1995) comprises 42 items in three-dimensional structure: depression (e.g., “I felt that I had lost interest in just about everything”), anxiety (e.g., “I feared that I would be ‘thrown’ by some trivial but unfamiliar task”), and stress (e.g., “I was in a state of nervous tension”). Participants responded on a four-point Likert scale ranging from 0 (did not apply to me at all) to 3 (applied to me very much, or most of the time). The DASS was adapted to Turkish by Akin and Çetin (2007) from the original version. The psychometric characteristics of the Turkish DASS are satisfactory (α = .90, .92, and .92, respectively). In the current research, Cronbach’s α coefficient values were .80 and .79 for positive affect and negative affect, respectively.

The Satisfaction with Life Scale (SWLS) by Diener, Emmons, Larsen, and Griffin (1985) comprises 5-items in uni-dimensional structure. Participants responded on a seven-point Likert scale ranging from 1 (strongly disagree) to 7 (strongly agree). A sample item is “The conditions of my life are excellent.” The scores that can be obtained in the SWLS vary between 5 and 35. Rising scores indicate that the level of life satisfaction also increased. The SWLS was adapted to Turkish by Durak, Senol-Durak, and Gençöz (2010) from the original version. The psychometric characteristics of the Turkish SWLS are satisfactory (α = .81; CFI = .99, IFI = .99, TLI = .98, SRMR = .020, and RMSEA = .043; Durak et al., 2010). In the current research, Cronbach’s α coefficient was .80.

PROCEDURE

Mediation was tested in IBM SPSS Statistics 22 using a macro developed by Hayes (PROCESS version 3.2; Hayes, 2018) in order to check the incremental validity of the Turkish OAS-2. This macro uses regression path analyses to estimate direct and indirect effects of a predictor variable on an independent variable. We first examined whether external shame (independent variable) was associated with subjective well-being (dependent variable) and depression, anxiety, and stress (DAS; mediators). Next, we examined whether the effect of external shame on subjective well-being (positive affect + life satisfaction – negative affect) could be mediated by psychological distress such as depression, anxiety, and stress. Bias-corrected 95% bootstrapped confidence intervals (CIs) of this indirect effect with 10 000 iterations were acquired. CIs that do not contain 0 indicate the presence of a mediation effect. This procedure is suggested for its accuracy in estimating confidence intervals and for control of type 1 error rates (MacKinnon, Lockwood, & Williams, 2004).

RESULTS

Findings for the regression pathways and the direct and indirect effects examined for mediation are presented in Figure 2 and Table 6.

Figure 2 shows the effect of external shame on subjective well-being through depression, anxiety, and stress. The total indirect effect is significant for external shame on subjective well-being (B = −1.33, SE = .12, 95% CI = −1.59 to −1.12). When depression, anxiety and stress are entered into the model as mediators, the direct effect of external shame on subjective well-being is significantly reduced, suggesting
According to the CFA conducted in Study 1, the OAS-2 was found to have acceptable goodness of fit indices. The CFA results indicated that all goodness of fit indices were at an acceptable level and that standardized factor loadings were significant. Therefore, the unidimensional structure of the Turkish OAS-2 was verified. Although the first version of the OAS involved a three-dimensional structure, the unidimensional structure was used in later studies (e.g., Gilbert & Miles, 2000; Matos et al., 2013). Matos and colleagues (2015), however, found that there was also a partial mediating role. Of the proposed mediators, the specific indirect effects are significant for depression ($B_{a1} = .92, SE = .07, 95\% CI = .78-.106$), anxiety ($B_{a2} = .78, SE = .07, 95\% CI = .64-.92$), and stress ($B_{a3} = .70, SE = .08, 95\% CI = .54-.86$). Consistent with expectations, the Turkish OAS-2 supported incremental validity and it predicted relevant constructs independent of other factors that have been shown to be statistically associated with these constructs.

**DISCUSSION**

The OAS-2 is a tool widely used in the world for measuring external shame. It has been adapted into various languages, and its validity and reliability have been previously examined. In this study, the psychometric properties of the OAS-2 in Turkish university students were evaluated. Three separate studies were conducted in line with this purpose, and results on external shame were obtained.

According to the CFA conducted in Study 1, the OAS-2 was found to have acceptable goodness of fit indices. The CFA results indicated that all goodness of fit indices were at an acceptable level and that standardized factor loadings were significant. Therefore, the unidimensional structure of the Turkish OAS-2 was verified. Although the first version of the OAS involved a three-dimensional structure, the unidimensional structure was used in later studies (e.g., Gilbert & Miles, 2000; Matos et al., 2013). Matos and colleagues (2015), however, found that there was also a partial mediating role. Of the proposed mediators, the specific indirect effects are significant for depression ($B_{a1} = .92, SE = .07, 95\% CI = .78-.106$), anxiety ($B_{a2} = .78, SE = .07, 95\% CI = .64-.92$), and stress ($B_{a3} = .70, SE = .08, 95\% CI = .54-.86$). Consistent with expectations, the Turkish OAS-2 supported incremental validity and it predicted relevant constructs independent of other factors that have been shown to be statistically associated with these constructs.

**Table 6**

| Path                        | Estimate | SE  | Lower 95% CI | Upper 95% CI |
|-----------------------------|----------|-----|--------------|--------------|
| Direct                      |          |     |              |              |
| OAS-2 → Depression          | .92      | .07 | .78          | 1.07         |
| OAS-2 → Anxiety             | .78      | .07 | .64          | .92          |
| OAS-2 → Stress              | .70      | .08 | .54          | .86          |
| OAS-2 → SWB                 | -.52     | .12 | -.76         | -.28         |
| Depression → SWB            | -.51     | .12 | -.75         | -.27         |
| Anxiety → SWB               | -.26     | .13 | -.51         | -.01         |
| Stress → SWB                | -.23     | .10 | -.42         | -.03         |
| Indirect effect             |          |     |              |              |
| OAS-2 → Depression → SWB    | -.47     | .12 | -.71         | -.24         |
| OAS-2 → Anxiety → SWB       | -.21     | .09 | -.39         | -.04         |
| OAS-2 → Stress → SWB        | -.16     | .07 | -.31         | -.03         |
| Total effect                | -1.33    | .12 | -1.59        | -1.12        |

*Note.* OAS-2 – Other as Shamer Scale-2; SWB – subjective well-being; SE – standard error; CI – confidence interval.
a unidimensional structure in the OAS-2 development study. Similarly, in the Italian sample (Saggino et al., 2017), the OAS-2’s unidimensional structure had good fit indices. Thus, the Turkish OAS-2 has both replicated its original unidimensional structure and verified the theoretical model of external shame.

In the configural and metric invariance analysis, the unidimensional structure proposed for gender yielded similar results in both females and males. This was consistent with the results showing that Portuguese adult females and males do not differ significantly and were partially invariant across boys and girls, who presented similar levels of shame in the long and short forms of the OAS (Matos et al., 2015; Vagos, da Silva, Brazao, Rijo, & Gilbert, 2016). These invariance results also parallel the Italian, university students’ version of the OAS-2 (Saggino et al., 2017). On the other hand, scalar and strict invariance analyses did not support the model in this present study, which, nevertheless, was also partially in parallel with the Italian version. Thus, the conceptualization of the external shame structure seems to be partially similar for both genders, as assessed by the OAS-2, given the results of configural and metric invariance.

No previous study in which OAS was examined using the Item Response Theory was encountered. There are two approaches: IRT and classical test theory (CTT). CTT has a major drawback. The summed scores are sequential, and statistical inferences based on sequence scores may be invalid because of the assumption that raw scores should be considered as a range scale (Hobart, Cano, Zajicek, & Thompson, 2007; Jafari, Bagheri, Ayatollahi, & Soltani, 2012). For example, CTT treats a five-point (never, seldom, sometimes, often, always) Likert-type scale as an interval scale, and scores five points between zero and four. However, the differences between the five points may not be equal. In contrast, the IRT uses a logistic equation to predict participants’ underlying capabilities and item challenges. Thus, the interval-level measurement of each participant’s external shame, and the interval-level measurements of item challenges, are determined (Chang, Lin, Gronholm, & Wu, 2018). The characteristics and validity of the items are examined in IRT, and the fitness of the items is determined. Therefore, according to the IRT results of the Turkish OAS-2, the item challenges and characteristics were appropriate. Researchers (Embretson & Reise, 2000; Reeve & Fayers, 2005) agree when using the IRT model with a five-point Likert-type scale, and 250 participants can produce a reasonable estimate, which indicates that the IRT results obtained are sufficient for the OAS-2.

Within the scope of Study 2, the criterion-related validity of the Turkish OAS-2 was discussed. The findings indicated that the OAS-2 was significant in the positive association with loneliness and in the negative association with subjective happiness and flourishing, which are indicators of wellbeing. The relevant literature suggested that shame is consistently and positively correlated with loneliness (e.g., Mereish & Poteat, 2015; Thoresen et al., 2018; Rostami & Jowkar, 2016) and has a negative relationship with well-being (e.g., Choma et al., 2009; Kelly & Carter, 2013). Therefore, this study’s results parallel the literature and support the criterion-related validity of the OAS-2.

Study 3 tested the incremental validity of the OAS-2, examining the mediation role of psychological distress in the relationship between external shame and well-being. The study found that external shame may increase depression, anxiety, and stress and that these increases may weaken an individual’s well-being. The literature supports this result (e.g., Castilho et al., 2017; Lee et al., 2016; Mereish & Poteat, 2015; Thoresen et al., 2018; Varghese, 2015). Kim and colleagues (2011) indicated that external shame may lead to psychological distress. Furthermore, external shame may also reduce well-being via psychological distress dimensions such as negative emotions, and feelings of helplessness and worthlessness. In addition, psychological distress being instrumental between shame and well-being may be explained by the fact that shamed individuals are passive against distress. In addition, it can be deduced that shame for mediation may make individuals vulnerable to psychological distresses, such as depression, anxiety, etc., and, hence, their well-being can be adversely affected (Clapton et al., 2018).

Finally, the internal, composite, and temporal reliability of the Turkish OAS-2 was examined. The reliability coefficient of OAS-2 in its original study was found to be .82. In addition, the OAS-2’s internal consistency was .89 for Italian university students (Saggino et al., 2017), .92 for Portuguese adolescents (Cunha et al., 2017), and .85 for children (Benevides et al., 2016). The reliability coefficients in this research are similar (.88, .89, and .89, Study 1, 2, and 3, respectively). However, the composite reliability of OAS-2 has not been previously discussed. CR in this study is sufficient, with .88. Temporal reliability was also studied for Portuguese adolescents (.73; Cunha et al., 2017) and Portuguese children (.55; Benevides et al., 2016). In this present research temporal reliability was found to be .86. All the reliability coefficients found in this study are observed to be similar to coefficients of the OAS-2 in other cultures. The findings also indicate that the criterion of being over .70 for reliability, as recommended by Nunnally (1978), is overcompensated. Therefore, it should be stated that the Turkish OAS-2 is a reliable measurement tool.

**LIMITATIONS**

In this study, where strong psychometric results of the Turkish OAS-2 have emerged, there were some
limitations. Firstly, the research was conducted with non-clinical, university students. Therefore, the findings should not be generalized to the clinical population. However, in future studies, the discriminant validity of the OAS-2 in clinical and non-clinical populations should be examined. Secondly, the Turkish OAS-2 is based on self-reported information. Because shame experiences may show partial bias, participants may perform differently on some items when self-reporting (Saggino et al., 2017). Therefore, self-report measures can be integrated with a neurophysiological markers analysis in order to provide a comprehensive assessment of shame (Vagos et al., 2016). Finally, the gender invariance of the OAS-2 was examined. However, subsequent studies, as Saggino and colleagues (2017) suggested, should address the measurement invariance of OAS-2 in different cultures.

CONCLUSIONS

The Turkish OAS-2 was shown to have validity and reliability, according to the findings of three different studies using advanced techniques such as measurement invariance, incremental validity, and item response theory. The results indicate that the Turkish OAS-2 can be used to measure external shame in a short time. Although other studies on the OAS-2’s measurement power are needed, this research supports the idea that external shame is a useful tool for evaluating individuals’ personality, psychological distress, and well-being. The OAS-2 cannot be used as a tool that gives a general evaluation of individuals’ personality, but to evaluate the particular predisposition of imaging negative beliefs towards the self in the mind of others.

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