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Removing user fees to improve access to caesarean delivery: a quasi-experimental evaluation in western Africa

Marion Ravit,1 Martine Audibert,2 Valéry Ridde,3,4 Myriam de Loenzien,1 Clémence Schantz,1 Alexandre Dumont1

ABSTRACT

Introduction Mali and Benin introduced a user fee exemption policy focused on caesarean sections in 2005 and 2009, respectively. The objective of this study is to assess the impact of this policy on service utilisation and neonatal outcomes. We focus specifically on whether the policy differentially impacts women by education level, zone of residence and wealth quintile of the household.

Methods We use a difference-in-differences approach using two other western African countries with no fee exemption policies as the comparison group (Cameroon and Nigeria). Data were extracted from Demographic and Health Surveys over four periods between the early 1990s and the early 2000s. We assess the impact of the policy on three outcomes: caesarean delivery, facility-based delivery and neonatal mortality.

Results We analyse 99,800 childbirths. The free caesarean policy had a positive impact on caesarean section rates (adjusted OR=1.36; 95% CI 1.11 to 1.66; P≤0.001), particularly in non-educated women (adjusted OR=2.71; 95% CI 1.70 to 4.32; P≤0.001), those living in rural areas (adjusted OR=2.02; 95% CI 1.48 to 2.76; P≤0.001) and women in the middle-class wealth index (adjusted OR=3.88; 95% CI 1.77 to 4.72; P≤0.001). The policy contributes to the increase in the proportion of facility-based delivery (adjusted OR=1.68; 95% CI 1.48 to 1.89; P≤0.001) and may also contribute to the decrease of neonatal mortality (adjusted OR=0.70; 95% CI 0.58 to 0.85; P≤0.001).

Conclusion This study is the first to evaluate the impact of a user fee exemption policy focused on caesarean sections on maternal and child health outcomes with robust methods. It provides evidence that eliminating fees for caesareans benefits both women and neonates in sub-Saharan countries.

INTRODUCTION

The number of maternal deaths across the world has decreased considerably since 1990, dropping from 390,155 to 275,288 in 2015. Almost half of these deaths occurred in sub-Saharan Africa (SSA), where the maternal mortality ratio (MMR) remains high (374.9 per 100,000) compared with the global MMR (195.7 per 100,000). One crucial issue that allows the reduction of maternal mortality is access to Emergency Obstetric and Neonatal Care. Approximately 15% of women are confronted with obstetric complications during their pregnancy, and between 3.6% and 6.5% of pregnant women should benefit from caesarean section (CS) delivery at birth in West Africa. The WHO suggests that at least 5% of pregnant women require a CS delivery.

Many women in low-income and middle-income countries (LMICs) delay their decision to seek care or give up going to a health facility due to fear of having to pay for excessive expenses. Household wealth directly
influences access to CS.9 The cost of a caesarean episode can be very high, up to 10 times higher than for a normal delivery, and could lead households to catastrophic expenses.7,10–13 The cost of a CS is also high for the health system. In 2008, the cost of the global needed CS was estimated to be approximately US$432 million.14

In recent years, the concept of universal health coverage (UHC), which posits that everybody should have access to health services without suffering from financial issues when paying their fees, has emerged.15 16 In September 2015, the Sustainable Development Goals were adopted by the United Nations, and a part of goal 3 consists of achieving UHC.17 A common goal of different UHC programme memes around the world is to remove financial barriers such as user fees, particularly for pregnant women and children, who require priority access to health services.18 19

User fee exemptions at the point of service are one reform that was implemented in LMICs to remove financial barriers in health services utilisation. Fee removal can concern a targeted intervention, such as CS, or a targeted population, such as pregnant women, depending on the country. Removing user fees for facility-based delivery (FBD) services increased the access to facility births, and to a much lesser extent, to caesarean sections in Ghana and Burkina Faso.20 One study, however, showed positive impact on FBD but failed to identify an effect on caesarean deliveries in Kenya, Ghana and Senegal.21 In contrast, user fee exemptions focused on CS, called the ‘free caesarean’ policy, only concerns women delivering by CS. In this case, women must pay for all other pregnancy-related care. Benin and Mali introduced such a policy in 2005 and 2009, respectively. A study showed that there is no evidence of a supplementary increase of CS rates after the implementation of the exemption policy in these two countries.22 We found no study assesses the impact of such reforms on health services utilisation with robust methods, using population data and comparison groups.

Our hypothesis is that this specific reform would increase access to CS and FBD and that this increase in utilisation of maternal health services would have positive repercussions on neonatal mortality. Some studies showed that an increase of CS rates, when these rates are below 5%–10%, could impact substantially neonatal mortality.23 Using long-term population data and a quasi-experimental approach, this study aims to assess the impact of a ‘free caesarean’ policy on access to caesarean and FBD and on early neonatal mortality in Mali and Benin. We focused specifically on whether this reform impacts women differentially by education level, zone of residence (rural vs urban) and wealth quintile of the household.

**DATA AND METHOD**

**Selection of countries**

To select comparison countries to evaluate impact of the free caesarean policy in Benin and Mali with a counterfactual, we first considered all sub-Saharan countries with Demographic and Health Surveys (DHS), which contains information on FBDs, CS deliveries and the neonatal health status of women who reported a birth during the last 3 or 5 years before the survey. In addition to Mali and Benin, we selected a total of 37 countries to form this sample. As four DHS surveys are available between the early 1990s and the early 2000s for Mali and Benin, we chose to select the 14 countries with at least four DHS surveys between the same period to use data in a panel manner and to be able to assess the impact of the reform with robust statistical analysis. We then concentrated on the six western and central African countries to study comparable countries. Among these countries, we excluded Côte d’Ivoire due to the situation of conflict that existed during the period of study, which can significantly affect the health system.24 Three additional countries that implemented a pregnancy-related fee exemption not specifically focused on CS were also excluded (Senegal, Niger and Ghana).25 Two countries did not experiment with any user fee exemptions for delivery between the first and the last available survey at a national level: Cameroon and Nigeria. These two countries were used as counterfactuals in our study (see figure 1).

**Description of intervention**

Mali introduced the user fee removal reform on 1 January 2005. This national policy concerns all CS in the public sector and covers the surgical procedure and preoperative examinations, the surgical kit and postoperative treatment (standardised set of products and medications) and hospitalisation. Structures receive 30 000 FCFA (US$50) for a simple caesarean and 42 000 FCFA (US$70) for a complicated caesarean, in addition to a surgical kit.26 27 In Benin, the user fee exemption policy was introduced on 1 April 2009 and concerns caesareans in all public and private hospitals that offer emergency obstetric care. Hospitals receive 100 000 FCFA (US$166) per caesarean. The policy covers check-up costs before medical intervention, drugs, kits, surgery, blood and hospitalisation for 7 days.28 In these two countries, the state is the principal financer of the reform.22

It is important to note that the level of implementation of the free caesarean policy in both countries is far from optimal and heterogeneous across countries. In Benin, the amount of the reimbursement is unique for all sites, while the cost of CS depends on the type of healthcare facility. Some authors reported that many pregnant women might have to pay a part of the caesarean fees not covered by the exemption policy.28 Because of the lack of official documents on the policy, there is no clear information for managers, health workers or even pregnant women on the real content of the reform. It induces different interpretation of the free package by health providers and a disappointment for women who have understood the caesarean was free in the whole country.29 In Malian health facilities, the introduction of free caesarean policy increased workload in public services, without financial compensation for health workers.20 It
puts more pressure on outdated equipment and leads to a deterioration of the quality of care. Furthermore, the standardised kits were often incomplete or sometimes not delivered in healthcare facilities during a long period of time. In a rural area of Mali, 91% of women still paid for their CS during the period 2008–2011.

**Data sources**

The four DHS surveys available for the included countries are distributed as follows: (1) Benin: 1996, 2001, 2006 and 2011–2012; (2) Mali: 1995–96, 2001, 2006 and 2012–13; (3) Cameroon: 1991, 1998, 2004 and 2011; and (4) Nigeria: 1990, 2003, 2008 and 2013.

We used data on women aged 15–49 years who spent the night before the interview in the surveyed household and who delivered a live-born child in the 3 years before the survey. If a woman delivered more than one child during this period, we selected the most recent birth only. We considered only the last birth for a woman to minimise recall bias regarding the history of previous pregnancies.

The DHS surveys are nationally representative household surveys with a large sample size (between 4500 and 38,500 households per survey for the four selected countries over the four periods). Data on household characteristics (demographic, socioeconomic and environmental conditions) and on the last pregnancy, including information on the use of maternal health services, were extracted from relevant questionnaires. During the interview, the woman was asked whether her child was born by CS. As recommended, and for more accuracy, cases of CS among live births that occurred at home were recoded as vaginal birth deliveries. Socioeconomic status was evaluated using principal component analysis as a relative wealth index that is based on household conditions and assets. We used the household wealth quintile available in each DHS survey.

**Measures**

We assessed the impact of the caesarean-related fee removal policy on three binary outcomes: CS, FBD and neonatal death (up to 28 days). CS is the primary outcome because we supposed that the policy implementation had a direct positive impact on women’s access to CS. FBD and neonatal mortality were chosen as secondary outcomes to assess if the policy helped improve access to healthcare facilities and child survival. An FBD was defined as a birth that took place in a public or private healthcare facility (hospital, health centre and clinic). Other births (taxi, car and home) were coded as non-FBD. In the case of FBD, we created a specific variable to study if the birth occurred in a hospital or in another category of healthcare facility. Neonatal mortality was defined as an infant death during the first 28 days of life (0–28 days). To assess the outcome of each included birth, we used information on infant mortality (age at death in days, months or years) reported by the interviewed mothers.

**Exposure variable**

We created a variable denominated ‘free caesarean’ (yes or no) to indicate if the birth occurred in a country that implemented a caesarean-related fee exemption policy (Mali and Benin) or one that did not (Cameroon and Nigeria) over the studied period. We identified four studied periods according to the time of each DHS: period 1 (1990–1996), period 2 (1998–2003), period 3 (2004–2008) and period 4 (2011–2013). We defined two time periods according to the year of implementation of the ‘free caesarean’ policy in Mali and Benin: before and after the policy implementation. Each birth occurring after January 2005 in Mali and after April 2009 in Benin corresponds to the period ‘after’. For the control group (women who give birth in a country that did not implement a caesarean-related fee exemption policy), we defined 1 January 2008 as the criteria for the ‘time’ variable because the first birth in the last survey for the Cameroon occurred in 2008. We classified each birth occurring...
after January 2008 as ‘after’ for the time variable in the group of country with no free caesarean policy.

**Statistical analysis**

The effect of the policy was assessed as the difference in the change in the primary outcome from the period before to the period after implementation between the two groups of countries (free or unfree caesarean).\(^3\)\(^4\)\(^5\) The additional increase in the risk of having a CS in the free caesarean group, relative to the increase in this risk in the other group, was estimated by the OR with 95% CIs for the interaction between groups (free and unfree caesarean) and time (before and after) provided by the logistics regression model. The model adjusted for variables selected a priori as potential effect modifiers for CS: zone of residence, maternal age, education level, wealth quintile of household, parity and multiple pregnancies. The model took into account sampling weight, clustering and strata. Secondary binary outcomes, related to either FBD or neonatal deaths, were analysed using the same method as for the primary outcome. We considered two different types of FBD: delivery at referral hospital and delivery at another type of healthcare facility (health post, health centre or private clinic). We supposed that the impact of a free caesarean policy would be more important for delivery at hospital because this is the place where the surgery is performed.

Furthermore, we assessed whether the free caesarean policy effects varied between educated and non-educated women, between those living in an urban and rural context and between household wealth quintiles. The free caesarean policy effects were reported separately for each stratum of the respective variable.

We performed sensitivity analysis by recoding the ‘group’ variable into three categories: Benin, Mali and control countries (Cameroon and Nigeria). The coefficient of interaction between group (three categories) and time in the regression model measures the effect of the policy in Benin, on the one hand, and in Mali, on the other hand, compared with the control group.

To use the differences-in-differences method, we made the assumption that changes in the proportion of CS, FBDs and neonatal mortality would not differ between the intervention and control groups without the presence of the policy. We visually checked this assumption by ensuring that trends in the predicted prevalence of the three outcomes, adjusted on variables cited above, are similar for the intervention and control areas before the implementation of the policy.

No imputation of missing data was performed. Tests were two tailed, and \(P<0.05\) was considered statistically significant. We managed the data using SPSS V.20, and analyses were performed using Stata V.13.0.

**RESULTS**

**Descriptive analysis**

The four DHS surveys available for each country allowed us to analyse 99,800 women who delivered a live child in the last 3 years before the interview, 46,362 women in the ‘free caesarean’ policy group and 53,438 women in the ‘no free caesarean’ policy group. The mode of delivery was available for 99.5% of the 99,800 women included (99,324 women), and the place of delivery was reported for 99,476 women (99.7%). There are no missing data for neonatal mortality, as we considered a child to be alive if the infant’s mortality was not reported. Table 1 presents the baseline characteristics of these women by country (ie, prior the implementation of intervention). We observed no major differences in these characteristics except for education level. At baseline, women in Benin and Mali are less educated than those of Cameroon and Nigeria.

Figure 2A illustrates the trend in the adjusted predicted prevalence of caesarean sections from period 1 to period 4 by each of the studied countries. We do not observe differences in the trend in adjusted predicted prevalence of caesarean section prior the implementation of the policy in the different countries. Similar patterns can be observed for FBD and neonatal mortality (see Figure 2B,C).

**Differences-in-differences analysis**

After adjusting on covariates (Table 2), there was a significant increase in the rate of CS in countries with a free caesarean policy compared with countries without such a policy, with an adjusted OR of 1.36 (95% CI 1.11 to 1.66; \(P\leq0.01\)) and an adjusted absolute risk difference of 1.0% (95% CI 0.6% to 1.4%; \(P\leq0.001\)). Furthermore, the caesarean rates increased significantly more rapidly in free caesarean countries for non-educated women (adjusted \(OR=2.71; 95\%\ CI 1.70 to 4.32; P\leq0.001\)), for women who live in rural areas (adjusted \(OR=2.02; 95\%\ CI 1.48 to 2.76; P\leq0.001\)), Q1 wealth index women (adjusted \(OR=2.02; 95\%\ CI 1.07 to 3.84; P\leq0.05\)) Q3 wealth index women (adjusted \(OR=3.88; 95\%\ CI 1.77 to 4.72; P\leq0.001\)) and Q4 wealth index women (adjusted \(OR=1.56; 95\%\ CI 1.05 to 2.33; P\leq0.05\)).

Table 3 shows that FBD rates (adjusted \(OR=1.68; 95\%\ CI 1.48 to 1.89; P\leq0.001\)) and particularly hospital-based delivery rates (adjusted \(OR=1.71; 95\%\ CI 1.49 to 1.95; P\leq0.001\)) increased significantly more rapidly in Mali and Benin compared with Cameroon and Nigeria. This positive effect of free caesarean policy on FBD is observed for every subgroup of women (education level, zone of residence and wealth).

As shown in Table 4, a free caesarean policy resulted in a significant decrease in neonatal mortality (adjusted \(OR=0.70; 95\%\ CI 0.58 to 0.85; P\leq0.001\)), especially in non-educated women (adjusted \(OR=0.75; 95\%\ CI 0.58 to 0.95; P\leq0.05\)), Q3 wealth index women (adjusted \(OR=0.55; 95\%\ CI 0.36 to 0.84; P\leq0.01\)) and Q4 wealth index women (adjusted \(OR=0.59; 95\%\ CI 0.38 to 0.91; P\leq0.05\)).

We observed differential effect of the policy according to the intervention country. We found that the impact of the policy on caesarean rates is significant in Mali (adjusted \(OR=1.89; 95\%\ CI 1.42 to 2.51; P\leq0.001\)) and...
Table 1  Baseline characteristics of pregnant women* by country

|             | Benin | Mali | Cameroon | Nigeria |
|-------------|-------|------|----------|---------|
| N           | 14397 | 14857| 8082     | 18136   |
| Age categories |       |      |          |         |
| <18 years   | 220   | 501  | 364      | 480     |
| 18–34 years | 11185 | 11015| 6438     | 13470   |
| 35 years and more | 2992 | 3341 | 1280     | 4186    |
| Parity      |       |      |          |         |
| Primiparous | 2712  | 2434 | 1815     | 3237    |
| Multiparous | 11685 | 12423| 6267     | 14899   |
| Multiple pregnancy | | | | |
| Simple      | 13974 | 14615| 7893     | 17800   |
| Multiple    | 423   | 242  | 189      | 336     |
| Zone of residence | | | | |
| Urban       | 4756  | 3938 | 3411     | 5427    |
| Rural       | 9641  | 10919| 4671     | 12709   |
| Education level | | | | |
| None        | 10837 | 12484| 2313     | 9403    |
| Primary or more | 3560 | 2373 | 5769     | 8733    |
| Wealth quintiles of households | | | | |
| Q1 poorest | 3369  | 2984 | 1702     | 4471    |
| Q2 poorer  | 3077  | 3017 | 1559     | 4144    |
| Q3 middle  | 3031  | 3169 | 1649     | 3256    |
| Q4 richer  | 2826  | 3139 | 1589     | 3217    |
| Q5 richest | 2094  | 2548 | 1583     | 3048    |

*Women who delivered a live-born child in the last 3 years before interview.

marginal in Benin (adjusted OR=1.25; 95% CI 0.99 to 1.56; P=0.055).

The impact of the policy was similar in both countries in terms of FBD (adjusted OR for Mali=2.09; 95% CI 1.79 to 2.45; P=0.001; adjusted OR for Benin=2.67; 95% CI 2.20 to 3.25; P=0.001) and neonatal mortality (adjusted OR for Mali=0.63; 95% CI 0.51 to 0.78; P=0.001; adjusted OR for Benin=0.76; 95% CI 0.57 to 0.98; P=0.039).

DISCUSSION

Statement of principal findings

Caesarean delivery was increased by 36% in Benin and Mali after the free caesarean policy was implemented. Non-educated women and those living in rural areas or in the poorest and middle-class households benefited the most from the reforms. Finally, the policy also resulted in a marked increase in facility births and a significant decrease in neonatal mortality, especially in disadvantaged groups, such as non-educated and rural women.

In this quasi-experimental study, we used counterfactual countries (Cameroon and Nigeria) to assess the impact of free caesarean policies in Benin and Mali. This design is recommended to evaluate the effects of fee exemption policies on maternal and neonatal health when randomised controlled studies are not possible. To our knowledge, this is the first study to assess the impact of a user fee exemption policy focused on CS with a robust method.

Our findings complement other studies of pregnancy-related user fee removal policies (but not specifically for CS) in SSA countries. All of these studies found an increase in the utilisation of maternal health services after the reform, especially in FBD. However, none of these authors, except Leone et al, found evidence of any positive effect on caesarean deliveries. The other study that evaluated the impact of the exemption policy in Mali and Benin using segmented regression analysis found no effect on CS rates after the implementation of the policy but they did not use comparison group. Our findings demonstrate that a user-fee removal policy focused on CS had a positive and marked impact on CS. A systematic review reported that the increase of caesarean delivery rates of up to 10% are associated with decreased neonatal mortality. The different studies failed to control for concurrent reduction in neonatal mortality, however, reflecting secular trends. Our study showed that the free caesarean policy reduced neonatal
mortality by 30%, relative to the simultaneous changes in countries without a user fee removal policy. However, these results should be interpreted with caution because neonatal mortality starts to decrease in the middle of 2000s, before the introduction of the free CS policy in Mali and Benin (see figure 2A). Using a quasi-experimental design, McKinnon et al. also showed a reduction in neonatal mortality by removing user fees for FBD services.

This study found that the abolition of user fees for CS improves access to CS, especially for vulnerable groups, such as non-educated women and those who live in rural areas. We assume that vulnerable pregnant women in Benin and Mali were less afraid of catastrophic expenditures when faced by complications after the free caesarean policy was implemented. For this reason, it was easier to make the decision to seek care. Furthermore, when women needed a CS, the decision was less conditioned by the cost of the care and they were more confident in the system. Even if the free policy was not totally effective, the monetary burden was less important for women after the reform.

There are several limitations to this study. First, quasi-experimental methods differ from experimental methods in that they are not based on randomised assignment of the intervention. If any other factors are present and affect the difference in trends between the two groups of countries, the estimation could be invalid or biased. To address this issue, we searched the literature for other major policies and reforms in maternal health occurring during the studied period in each included country (see online supplementary file 1). We could assume that the introduction of health insurance schemes in Mali and Nigeria improved pregnant women’s access to the maternal health services. In Mali, however, despite the introduction of health insurance targeted to poor people, approximately 80% of the population did not have medical coverage in 2012. In Nigeria, only approximately 3.5% of the population was enrolled (5.3 million persons) in 2010. In Mali, a study of rural areas showed that the referral evacuation system had a positive impact on CS rates. This system could partially explain our results through improvements in transportation and in the quality of emergency obstetric care. In Nigeria, the fee exemption policy on maternal and newborn care concerned 14–36 states and was implemented heterogeneously according to the state, but we did not find information on what the exemption covers or its effectiveness. Moreover, in the early 2010s, most primary healthcare facilities are unable to adequately provide basic emergency obstetric care services or meet the demand for obstetric care. We can assume that these local interventions did not have a major impact on maternal health service utilisation. We assumed that other policies in Benin, Cameroon and Nigeria did not influence our results because they were introduced in 2012 at the end of our study period.

Second, the selection of comparison countries (Cameroon and Nigeria) can be considered arbitrary. In fact, we chose to exclude countries such as Senegal, Niger and Ghana because they implemented a user fee exemption policy that was not particularly focused on CS. If we had taken them into account as intervention countries, it is likely that the effect size of the policy would have been underestimated. Furthermore, the choice of timing to classify births as prepolicy or postpolicy for comparison countries (1 January 2008) can be also considered arbitrary. If we had chosen a different criterion for the time variable, we can suppose that the study results would have been different.

Third, DHS collects information on live births only. Therefore, all stillbirths were omitted in the analyses. The rate of stillbirths is high in western African countries, and there is evidence that the CS rate and intrapartum stillbirth rate are correlated. In this study, however, we could not show any impact of the free caesarean policy on stillbirth.

Fourth, the period ‘after’ the policy is different between
Table 2  Effects of ‘free caesarean’ policy on caesarean delivery

|                                | Countries with no ‘free caesarean’ policy | Countries with ‘free caesarean’ policy | Effect of ‘free caesarean’ policy |
|--------------------------------|------------------------------------------|----------------------------------------|----------------------------------|
|                                | Before After Diff. | Before After Diff. | Adjusted absolute risk difference (95% CI) | Adjusted ORs (95% CI) |
|-------------------------------|------------------|------------------|---------------------------------|-------------------|
| Total                         | 26081 27 104 0.5 | 29037 17 102 1.4 | 1.0 (0.6 to 1.4)*** | 1.36 (1.11 to 1.66)** |
| Countries by level of education |                               |                               |                                |                    |
| Educated                      | 14 413 15878 1.0 | 5884 3728 1.0 | 0.6 (−0.4 to 1.6) | 0.97 (0.76 to 1.25) |
| Non-educated                  | 11 668 11226 0.5 | 23 153 13 374 1.4 | 1.5 (1.1 to 1.9)*** | 2.71 (1.70 to 4.32)** |
| Countries by zone of residence |                               |                               |                                |                    |
| Urban                         | 8779 9070 0.3 | 8625 5400 1.4 | 1.0 (0.0 to 2.0)§ | 1.05 (0.81 to 1.38) |
| Rural                         | 17 302 18 034 1.7 | 20 412 11 702 1.0 | 1.1 (0.7 to 1.5)*** | 2.02 (1.48 to 2.76)** |
| Countries by wealth quintiles of households |                               |                               |                                |                    |
| Q1 poorest                    | 3151 6154 0.0 | 6310 3490 1.0 | 0.7 (0.4 to 1.3)** | 1.52 (0.89 to 2.6) |
| Q2 poorer                     | 5675 6342 0.0 | 6051 3580 1.0 | 1.0 (0.4 to 1.6)*** | 2.02 (1.07 to 3.84)* |
| Q3 middle                     | 4884 5525 0.0 | 6163 3534 1.0 | 0.8 (0.4 to 1.5)** | 1.36 (0.88 to 2.1) |
| Q4 richer                     | 4772 4945 0.0 | 5910 3415 1.0 | 0.8 (0.4 to 1.5)** | 1.36 (0.88 to 2.1) |
| Q5 richest                    | 4599 4138 0.0 | 4603 3083 1.0 | 0.8 (0.4 to 1.5)** | 1.36 (0.88 to 2.1) |

Source: DHS surveys (http://dhsprogramme.com/).
*P≤0.05; **P≤0.01; ***P≤0.001.
†The adjusted absolute risk difference represents adjusted differences between group-specific changes over time (adjusted on age, education, wealth quintile of household, parity, multiple pregnancy and took into account sampling weight).
‡The adjusted ORs for the interaction between groups (free caesarean or not) and time (before vs after the implementation of the policy) were estimated using multivariate logistic regression models adjusted on maternal age, zone of residence, education, wealth quintile of household, parity and multiple pregnancy and took into account sampling weight, clustering and strata.
§Marginal level of significance (0.05 < m < 0.10).
Countries with ‘free caesarean’ policy: Benin and Mali. Countries with no ‘free caesarean’ policy: Cameroon and Nigeria.
CS, caesarean section.
| Table 3 | Effects of ‘free caesarean’ policy on facility-based delivery |
|-------------------|-----------------------------|-----------------------------|-----------------------------|
| Countries with no ‘free caesarean’ policy & Countries with ‘free caesarean’ policy & Effect of ‘free caesarean’ policy |
| Before & After & Diff. & Before & After & Diff. & Adjusted absolute risk difference (95% CI)† & Adjusted ORs (95% CI)‡ |
| All pregnant women & 26129 & 11.593 & 29097 & 17.088 & 8.7 (7.7 to 9.7)*** & 1.68 (1.48 to 1.89)*** |
| FBD rate (%) & 42.4 & 42.7 & 0.3 & 57.1 & 66.8 & 9.7 (8.1 to 11.3)*** & 1.71 (1.49 to 1.95)*** |
| FBD & 11078 & 11.593 & 16607 & 11.412 & 9.2 (7.8 to 10.6)*** & 1.59 (1.36 to 1.87)*** |
| Hospital delivery rate (%) & 60.5 & 62.5 & 2.0 & 20.4 & 31.3 & 10.9 & 6.4 (4.6 to 6.2)*** & 2.14 (1.70 to 2.70)*** |
| Women by level of education & & & & & & & & & |
| Educated & 14456 & 15.946 & 5895 & 37.25 & 64.0 & 63.0 & −1.0 & 9.2 (7.8 to 10.6)*** & 1.59 (1.36 to 1.87)*** |
| FBD rate (%) & 67.8 & 70.0 & 2.2 & 79.8 & 89.3 & 9.5 & 6.4 (4.6 to 6.2)*** & 2.14 (1.70 to 2.70)*** |
| FBD & 117329 & 11.216 & 23202 & 13.363 & 5.0 (3.0 to 7.0)*** & 1.62 (1.35 to 1.96)*** |
| Non-educated & 115673 & 11.946 & 16607 & 11.412 & 50.5 & 60.5 & 10.0 & 9.2 (7.8 to 10.6)*** & 1.59 (1.36 to 1.87)*** |
| FBD rate (%) & 15.6 & 13.8 & −1.8 & 5895 & 37.25 & 64.0 & 63.0 & −1.0 & 9.2 (7.8 to 10.6)*** & 1.59 (1.36 to 1.87)*** |
| Women by zone of residence & & & & & & & & & |
| Urban & 8800 & 9142 & 8643 & 5395 & 67.8 & 70.0 & 2.2 & 79.8 & 89.3 & 9.5 & 6.4 (4.6 to 6.2)*** & 2.14 (1.70 to 2.70)*** |
| FBD rate (%) & 67.8 & 70.0 & 2.2 & 79.8 & 89.3 & 9.5 & 6.4 (4.6 to 6.2)*** & 2.14 (1.70 to 2.70)*** |
| Rural & 17329 & 18020 & 20454 & 11.693 & 29.5 & 29.0 & −0.5 & 47.5 & 56.4 & 8.9 & 9.7 (8.3 to 11.1)*** & 1.62 (1.40 to 1.87)*** |
| FBD rate (%) & 29.5 & 29.0 & −0.5 & 47.5 & 56.4 & 8.9 & 9.7 (8.3 to 11.1)*** & 1.62 (1.40 to 1.87)*** |
| Women by wealth quintiles of households & & & & & & & & & |
| Q1 poorest & 6155 & 6141 & 6322 & 3487 & 15.1 & 9.2 & −5.9 & 37.2 & 50.2 & 13.0 & 16.8 (14.6 to 19.0)*** & 2.94 (2.34 to 3.71)*** |
| FBD rate (%) & 5681 & 6338 & 6061 & 3577 & 24.3 & 26.8 & 2.5 & 46.2 & 56.1 & 9.9 & 9.3 (6.9 to 11.7)*** & 1.47 (1.21 to 1.78)*** |
| Q2 poorer & 4896 & 5530 & 6173 & 3532 & 41.6 & 46.9 & 5.3 & 53.1 & 60.7 & 7.6 & 6.0 (3.3 to 8.7)*** & 1.31 (1.09 to 1.59)* |
| FBD rate (%) & 41.6 & 46.9 & 5.3 & 53.1 & 60.7 & 7.6 & 6.0 (3.3 to 8.7)*** & 1.31 (1.09 to 1.59)* |
| Q3 middle & 4781 & 4973 & 5928 & 3412 & 62.0 & 64.5 & 2.5 & 67.9 & 76.3 & 8.4 & 8.2 (5.7 to 10.7)*** & 1.66 (1.33 to 2.08)*** |
| FBD rate (%) & 62.0 & 64.5 & 2.5 & 67.9 & 76.3 & 8.4 & 8.2 (5.7 to 10.7)*** & 1.66 (1.33 to 2.08)*** |
| Q4 richer & 4616 & 4180 & 4613 & 3080 & 81.7 & 84.5 & 2.8 & 90.1 & 94.4 & 4.3 & 3.3 (1.3 to 5.3)** & 2.03 (1.44 to 2.85)** |
| FBD rate (%) & 81.7 & 84.5 & 2.8 & 90.1 & 94.4 & 4.3 & 3.3 (1.3 to 5.3)** & 2.03 (1.44 to 2.85)** |

Source: DHS surveys (http://dhsprogramme.com/).

†The adjusted absolute risk difference represents adjusted differences between group-specific changes over time (adjusted on age, education, wealth quintile of household, parity, multiple pregnancy and took into account sampling weight).

‡The adjusted ORs for the interaction between groups (‘free caesarean or not’) and time (before vs after the implementation of the policy) were estimated using multivariate logistic regression models adjusted on maternal age, zone of residence, education, wealth quintile of household, parity, multiple pregnancy and took into account sampling weight, clustering and strata.

§Marginal level of significance (0.05 < m < 0.10).

Countries with ‘free caesarean’ policy: Benin and Mali. Countries with no ‘free caesarean’ policy: Cameroon and Nigeria.

FBD, facility-based delivery.
Table 4  Effects of ‘free caesarean’ policy on neonatal mortality

| Countries with no ‘free caesarean’ policy | Countries with ‘free caesarean’ policy | Effect of ‘free caesarean’ policy |
|----------------------------------------|--------------------------------------|---------------------------------|
|                                        | Before  | After  | Diff.  | Before  | After  | Diff.  | Adjusted absolute risk difference (95% CI)† | Adjusted ORs (95% CI)‡ |
| All live births                        | 26218  | 27220  | 0.3    | 29254  | 17108  | 0.5    | −0.8 (−1.2 to −0.4)*** 0.70 (0.58 to 0.85)*** |
| Death per 100 births                   | 2.5     | 2.8     | 0.3    | 2.8     | 2.3     | 0.5    | −0.8 (−1.2 to −0.4)*** 0.70 (0.58 to 0.85)*** |
| By level of education of the mother    |         |         |        |         |         |        |                                                |
| Educated                               | 14502  | 15978  | 0.4    | 5933   | 3731   | 0.2    | −0.7 (−1.5 to 0.1) 0.72 (0.51 to 1.03) §       |
| Early neonatal death rate (%)          | 2.3     | 2.7     | 0.4    | 2.5     | 2.3     | 0.2    | −0.7 (−1.5 to 0.1) 0.72 (0.51 to 1.03) §       |
| Non-educated                           | 11716  | 11242  | 0.1    | 23321  | 13377  | 0.6    | −0.7 (−1.3 to −0.1) 0.74 (0.58 to 0.95)*       |
| Early neonatal death rate (%)          | 2.7     | 2.8     | 0.1    | 2.9     | 2.3     | 0.6    | −0.7 (−1.3 to −0.1) 0.74 (0.58 to 0.95)*       |
| By zone of residence of the mother**   |         |         |        |         |         |        |                                                |
| Urban                                  | 8838   | 9159   | 0.2    | 8694   | 5404   | 0.3    | −0.9 (−1.5 to −0.3)** 0.65 (0.46 to 0.92)*     |
| Early neonatal death rate (%)          | 2.3     | 2.5     | 0.2    | 2.5     | 2.2     | 0.3    | −0.9 (−1.5 to −0.3)** 0.65 (0.46 to 0.92)*     |
| Rural                                  | 17380  | 18061  | 0.7    | 20560  | 11704  | 0.8    | −0.8 (−1.4 to −0.2)** 0.73 (0.58 to 0.91)**    |
| Early neonatal death rate (%)          | 2.6     | 2.9     | 0.3    | 3.0     | 2.3     | 0.7    | −0.8 (−1.4 to −0.2)** 0.73 (0.58 to 0.91)**    |
| By wealth quintiles of household       |         |         |        |         |         |        |                                                |
| Q1 poorest                             | 6173   | 6155   | 0.2    | 6353   | 3491   | 0.7    | −0.3 (−1.1 to 0.5) 0.87 (0.60 to 1.26)         |
| Early neonatal death rate (%)          | 2.3     | 2.6     | 0.3    | 2.8     | 2.8     | 0.0    | −0.3 (−1.1 to 0.5) 0.87 (0.60 to 1.26)         |
| Q2 poorer                              | 5703   | 6350   | 0.0    | 6094   | 3580   | 0.5    | −0.6 (−1.6 to 0.4) 0.80 (0.54 to 1.18)         |
| Early neonatal death rate (%)          | 2.9     | 2.9     | 0.0    | 2.9     | 2.4     | 0.5    | −0.6 (−1.6 to 0.4) 0.80 (0.54 to 1.18)         |
| Q3 middle                              | 4905   | 5545   | 0.0    | 6200   | 3536   | 0.0    | −1.4 (−2.0 to −0.8)** 0.55 (0.36 to 0.84)**    |
| Early neonatal death rate (%)          | 2.6     | 2.9     | 0.3    | 3.0     | 1.9     | 1.1    | −1.4 (−2.0 to −0.8)** 0.55 (0.36 to 0.84)**    |
| Q4 richer                              | 4806   | 4985   | 0.0    | 5965   | 3416   | 0.7    | −1.1 (−2.0 to −0.2)** 0.55 (0.36 to 0.84)**    |
| Early neonatal death rate (%)          | 2.4     | 2.9     | 0.5    | 2.7     | 2.1     | 0.6    | −1.1 (−2.0 to −0.2)** 0.55 (0.36 to 0.84)**    |
| Q5 richest                             | 4631   | 4185   | 0.4    | 4642   | 3085   | 0.4    | −0.7 (−1.7 to 0.3) 0.71 (0.43 to 1.16)         |

Source: DHS surveys (http://dhsprogramme.com/).

*P<0.05; **P<0.01; ***P<0.001.
†The adjusted absolute risk difference represents adjusted differences between group-specific changes over time (adjusted on age, education, wealth quintile of household, parity, multiple pregnancy and took into account sampling weight).
‡The adjusted ORs for the interaction between groups (free caesarean or not) and time (before vs after the implementation of the policy) were estimated using multivariate logistic regression models adjusted on maternal age, zone of residence, education, wealth quintile of household, parity, multiple pregnancy and took into account sampling weight, clustering and strata.
§Marginal level of significance (0.05<m<0.10).
Countries with ‘free caesarean’ policy: Benin and Mali Countries with no ‘free caesarean’ policy: Cameroon and Nigeria.
both intervention countries: 8 years in Mali and 3 years in Benin. This could partly explain why the impact of the policy on caesarean rates in statistically significant in Mali and marginally significant in Benin. Finally, this impact study did not take into account the level of implementation of the free caesarean policy in both countries. As we mentioned previously, the implementation of the policy is not optimal in Benin and Mali. The impact of the policy might have been more important than that observed if CSs were really free for every woman.

We may wonder whether the user fee removal policy benefits the poorest women, those who could not afford a caesarean intervention before the policy’s introduction. This study showed that the effect on caesarean rates was more important for women in the middle and upper-middle wealth quintiles. Generally, only the poorest and the richest classes are studied, and the subject of the emerging middle classes in western Africa is not yet well developed in the literature. Further research is needed to understand the effect of user fee removal reforms in this subgroup. Further analyses of DHS are also required to examine if the free caesarean policy reduces inequalities in coverage and health outcomes.16

CONCLUSION
This study provides evidence that pregnant women’s access to facility-based and caesarean deliveries was improved in countries where free caesarean policies were implemented in the 2000s. It confirms the idea that user fees are one of the barriers preventing access to maternal health services, especially for CS. A free caesarean policy is entirely within the concept of UHC, because it is an effective way to provide CS to women in need. It seems important that the governments of Mali and Benin continue to finance this exemption policy and provide reimbursements, complete kits, a workforce and motivated health workers to health facilities in a timely fashion.

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