RESEARCH ARTICLE

Clarithromycin use and the risk of mortality and cardiovascular events: A systematic review and meta-analysis

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Abstract

Background

Although studies reported increased cardiovascular (CV) risks in patients treated with macrolides, the risks remain controversial among clarithromycin (CLR) users. We aimed to summarize the association between CLR use and the risks of mortality and CV events.

Methods

We searched PubMed, EMBASE, Web of Science, and the Cochrane Library for randomized controlled trials (RCTs) and observational studies with population exposed to CLR published until December 31st, 2018. These studies reported either all-cause mortality (primary outcome) or CV adverse events (secondary outcomes) based on multivariate models. Effect measures were synthesized by study design and follow-up duration (long-term, \( \geq 1 \) year; short-term, \( \leq 3 \) months; and immediate, \( \leq 2 \) weeks). This study has been registered on PROSPERO (ID: CRD42018089605).

Results

This meta-analysis included 13 studies (3 RCTs and 10 observational studies) and 8,351,815 subjects (1,124,672 cases and 7,227,143 controls). Overall, CLR use was not associated with increased long-term all-cause mortality (pooled rate ratio RR = 1.09, 95\% CI = 0.91–1.32), either among patients with or without comorbidities of cardiovascular diseases. Comparing CLR users to placebo, there is no additional risks of cardiac mortality (pooled RR = 1.03, 95\% CI = 0.53–2.01), acute myocardial infarction (pooled RR = 1.29, 95\% CI = 0.98–1.68), and arrhythmia (pooled RR = 0.90, 95\% CI = 0.62–1.32).

Conclusions

Our findings suggested no significant association between CLR use and subsequent long-term all-cause mortality, regardless having comorbidity of cardiovascular diseases or not.
Further RCTs investigating the short-term CV risks of CLR use compared to alternative antibiotics are warranted, particularly in high-risk populations.

Introduction

According to the 2011 statistics of outpatient insurance database, there are more than 260 million antibiotic prescriptions per year in the United States. Among them, macrolides (e.g., clarithromycin (CLR), erythromycin, azithromycin, roxithromycin, etc.) are the most common antibiotics. CLR has better bioavailability, gastrointestinal tolerance, and bactericidal activity than other macrolides, supporting its broad use from respiratory to gastrointestinal tract infections, as well as *Helicobacter pylori* infection. Previous meta-analysis suggested the use of macrolides being associated with increased risks of arrhythmia or QT interval prolongation through multivariate models, possibly due to electrophysiological side effects.

However, results from epidemiological studies using multivariate models regarding the cardiovascular (CV) risks associated with CLR use remained controversial because of different study designs and follow-up durations. One population-based study with a short-term follow-up period revealed a 2.2 times increased risk of arrhythmia among CLR users; whereas the study published by Chou et al. found no significant association between CLR use and the risk of arrhythmia within 30 days. Similarly, one large cohort study with a long-term follow-up period revealed an increased risk of acute myocardial infarction (MI) associated with CLR use, but the other studies didn’t.

In 2015, the CLARICOR trial showed an increased risk of all-cause mortality comparing CLR use versus placebo among patients with stable coronary artery disease (CAD). According to this single trial, the US Food and Drug Administration issued a warning in February 2018 to remind cautious use of CLR in CAD patients. At the same time, a meta-analysis of macrolides reported CLR users being associated with a 59% higher risk of acute MI compared to non-macrolide users. However, that study did not consider the differential effects in patients with various underlying diseases and follow-up durations.

In physicians’ perspective, the risks of CLR use should be compared with alternative antibiotics use, but not placebo. Knowing the risks of CLR use among populations with different diseases or follow-up durations is crucial for clinical decisions and patient selection. Therefore, a meta-analysis addressing these issues would bring clinical values. The objective of this systematic review and meta-analysis is to summarize the association between CLR use and the short-term and long-term all-cause and cardiovascular mortality in different study populations.

Materials and methods

Protocol and registration

We conducted the protocol of this study according to the Preferred Reporting Items for Systematic review and Meta-Analysis Protocols (PRISMA-P) 2015 statement. The present study has been registered with the International Prospective Registry of Systemic Reviews (PROSPERO CRD42018089605).

Data sources, searches and selection criteria

We searched publications in humans from databases of PubMed, EMBASE, Web of Science, and the Cochrane Library until December 31st, 2018. A manual search of the reference lists
Data extraction

Data from each enrolled study were extracted by two independent reviewers (CHY, PHC) to ensure quality and accuracy. Extracted variables included: author, journal, publication year, country, study design, sample size, underlying disease of subjects, exposure doses and durations of medication/placebo use, names of the comparator, follow-up durations, types of outcomes, and effect measures. Any disagreement in data extraction was discussed with a third reviewer until consensus was reached.

All-cause mortality was considered as primary outcome of interest because it is the most commonly reported endpoint. Acute myocardial infarction (AMI), cardiac mortality, and arrhythmia were considered as secondary outcomes. The definitions of arrhythmia in those enrolled studies consisted of a composite of atrioventricular conduction disorders, tachycardia and bradycardia accordingly. Follow-up duration was determined between the first date of antibiotic prescription and occurrence of outcome or censoring. One previous study used three months as the cut-off threshold for follow-up duration after CLR use.[11] Besides, the follow-up durations of the enrolled 13 study populations ranged from 5 days to 3 months, or 1 to 3 years. Thus, we divided the follow-up time into short-term (<3 months) and long-term (≥1 year). The immediate follow-up duration was defined as ≤2 weeks, since clarithromycin was rarely used for more than 2 weeks. The pooled effect estimates were calculated accordingly.

Statistical analysis

Random effects models [12] were chosen a priori to calculate summary effect estimates, weighting for inverse variance separately for RCTs and observational studies. For studies reporting multiple types of prescriptions for CLR, we chose a one-time prescription to align with the study results. We used rate ratio (RR) as the main pooled effect estimates, considering the time-to-event effect; the alternative was risk ratio if rate ratio was not available. For the estimates by odds ratio or hazard ratio, we would reach out to the authors for raw data or convert them to rate ratio estimates using the following equations [13]:

$$RR = \frac{OR}{(1 - r) + (r \times OR)}; \quad RR = \frac{1 - e^{HR \times LN(1-r)}}{r}$$

where OR is odds ratio; LN is natural logarithm formula; RR is rate ratio; HR is hazard ratio; and r is the outcome event rate for the reference group.

We estimated the pooled risk of the primary outcome of interest (i.e., long-term all-cause mortality) based on the multivariate models in those studies, followed by the estimates in different subgroups (e.g., placebo or alternative antibiotics and with or without comorbidities).
and follow-up durations. The pooled risks of secondary outcomes were further summarized by different follow-up durations and types of studies (i.e., RCT and observational studies) as well.

**Quality assessment**

Quality assessment was performed by two reviewers independently. The Cochrane Collaboration’s tool [14] was used to assess the risk of bias in RCTs, including five metrics (six items): (1) adequacy of randomization, (2) allocation concealment, (3) blinding, (4) completeness of outcome data, and (5) selective reporting. Each item was graded as low, high, and unclear risk of bias. Studies without any item of high risk of bias were considered as high quality. For cohort and case-control studies, we used the Newcastle-Ottawa Quality Assessment Scale (NOS) [15], with a full score of 9, containing three subsections (9 items): (1) selection; (2) comparability; and (3) exposure (for case-control study) or outcome (for cohort studies). Studies with scores equal or greater than 6 points were considered as high quality.

Potential publication bias was evaluated by the Egger’s regression test. The $I^2$ test and Q statistics were also calculated to assess heterogeneity between studies. An $I^2$ higher than 50% or $p$-value of the Q statistic less than 0.05 were considered as large heterogeneity.[16, 17] Leave-one-out sensitivity analyses were performed for effect estimates with large heterogeneity. The subgroup analyses were conducted by study types, quality of studies (NOS score), and effect measures (HR or not). All analyses were conducted using R version 3.6.1 and Stata version 15.0 (Stata Corporation, Texas, USA).

**Results**

Out of the 3,171 records from the initial screening, after excluding those not fit the inclusion criteria, 79 articles were reviewed for the full-text. The searching algorithm was presented in detail in the Fig 1.

This meta-analysis included three RCTs [8, 11, 18] and ten observational studies (nine cohorts [4–7, 19–23] and one nested case-control study [24]), with a total of 8,351,815 subjects (1,124,672 cases and 7,227,143 controls), as shown in Table 1. In RCTs, all participants had CAD before randomization, instead of acute bacterial infection, and these RCTs chose placebo as controls rather than alternative antibiotics. In ten observational studies, the study participants, either from the general population (8 studies) or having chronic comorbidities (2 studies, ischemic heart disease and chronic obstructive pulmonary disease), were indicated for antibiotic treatment for infectious diseases (e.g., pneumonia or H. pylori infection).[7, 22] Thus, they were compared to active comparator antibiotics, instead of placebo. Regarding to follow-up durations, short-term and long-term outcomes were reported in 7 and 5 observational studies, respectively. RCTs only reported long-term outcomes.

**Primary analysis**

Meta-analysis combining RCTs and observational studies showed that CLR use was not associated with long-term all-cause mortality (pooled rate ratio RR = 1.09, 95% CI = 0.91–1.32) (Fig 2). Comparing CLR with other comparators, no higher risk of long-term all-cause mortality was found (pooled RR = 1.06, 95% CI = 0.94–1.20), regardless of patients with or without comorbidity of cardiovascular diseases (Table 2). However, with limited evidence (study number, N = 2), among patients with CAD, CLR use was associated with a 24% increased rate of all-cause mortality (pooled RR = 1.24, 95% CI = 1.04–1.48), compared to placebo. The short-term risk of all-cause mortality, yet with limited evidence (N = 2) as well, was associated with CLR use comparing to alternative antibiotics, among patients without comorbidity (pooled RR = 1.63, 95% CI = 1.13–2.37).
**Secondary analysis**

There was no significant increased risk of any secondary outcomes for CLR use. Specifically, meta-analysis of three RCT's showed no increased risk of AMI comparing CLR versus placebo users (pooled RR = 0.63, 95% CI = 0.28–1.43) (Fig 3A). For observational studies, the summary estimate of five studies yielded no significant association between CLR use and increased risk of AMI after at least 1 year of follow-up (pooled RR = 1.29, 95% CI = 0.98–1.68) (Fig 3B). In terms of short-term cardiac mortality and arrhythmias, the pooled estimates of observational studies revealed a rate ratio of 1.03 (95% CI = 0.53–2.01) and a risk ratio of 0.90 (95% CI = 0.62–1.32), respectively, comparing CLR use versus alternative antibiotics use (Fig 3C and 3D). In subgroup analysis, CLR use was not associated with immediate risk of cardiac
### Table 1. Basic characteristics of 13 studies included in this meta-analysis.

| Studies            | Country                | Sample size (Trt/ctrl) | Comparator antibiotic | Comorbidities of study population | Outcomes               | Effect measures | Point estimates (95%CI) | Follow-up duration |
|--------------------|------------------------|------------------------|-----------------------|-----------------------------------|------------------------|------------------|-------------------------|---------------------|
| *RCT*              |                        |                        |                       |                                   |                        |                  |                         |                     |
| Sinisalo et al. (2002) | Finland               | 74/74                  | Placebo               | Acute non–Q-wave infarction or unstable angina | AMI                    | Risk ratio       | 0.36 (0.14–0.94)        | 1.5                 |
| Berg et al. (2005)  | Netherlands            | 238/235                | Placebo               | Before CABG surgery              | ACM AMI                | Rate ratio*      | 1.10 (0.45, 2.59)       | 2                   |
| Winkel et al. (2015) | Denmark                | 2,172/2,200            | Placebo               | Stable coronary heart disease     | ACM AMI                | Rate ratio**     | 1.25 (1.04, 1.49)       | 3                   |
| *Observational Study*  |                        |                        |                       |                                   |                        |                  |                         |                     |
| Andersen et al. (2010) | Denmark               | 1,205/437              | Non-CLR antibiotics   | Ischemic heart disease           | ACM                    | Rate ratio       | 1.07 (0.90, 1.26)       | 1                   |
| Hutson et al. (2012) | Canada                 | 59/295                 | AZM                   | General population (aged > 65 years) | AA                     | Risk ratio       | 0.65 (0.35, 1.23)       | 30                  |
| Schembri et al. (2013) | UK                    | (COPD) 281/1,062 (GAP) 980/651 | Non-CLR antibiotics COPD and CAP | (COPD) ACM AMI | Rate ratio** Rate ratio** | 1.15 (0.90, 1.49) 1.56 (0.81, 3.03) | 1 1 |
| Svenström et al. (2014) | Denmark               | 160,297/4,355,309      | Penicillin V          | General population               | CD                     | Rate ratio       | 1.76 (1.08, 2.85) 1.06 (0.62, 1.82) | 7 37 |
| Chou et al. (2015) | Taiwan                 | 393,243/1,102,358      | AMC                   | General population               | CD CD AA               | Rate ratio* Rate ratio* | 0.49 (0.33, 0.71) 0.41 (0.29, 0.56) 0.72 (0.53, 0.96) | 14 30 |
| Wong et al. (2016) | China                  | 90,411/186,888         | AMX                   | General population               | ACM CD AA              | Rate ratio Rate ratio | 1.97 (1.83, 2.11) 1.67 (1.36, 2.06) 0.94 (0.59, 1.51) | 14 14 |
| Mosholder et al. (2017) | UK                    | 287,748/267,729        | Doxycycline           | General population               | ACM AMI                | Rate ratio** Rate ratio** | 0.84 (0.80, 0.86) 0.99 (0.84, 1.18) | 1 1 |
| Inghammer et al. (2017) | Denmark               | 187,887/751,543        | Penicillin V          | General population               | CD ACM                 | Rate ratio Rate ratio | 1.66 (0.98, 2.79) 1.06 (0.98, 1.16) | 7 1 |
| Sutton et al. (2017) | US                     | 38,133/283,743         | AZM                   | General population               | CD                     | Rate ratio*       | 0.96 (0.38, 2.41)       | 5                   |

(Continued)
mortality (pooled RR = 1.19, 95% CI = 0.69–2.07) (Fig 3E), with a large heterogeneity ($I^2 = 87.9\%, P < 0.001$).

Quality assessment and publication bias

The detailed quality assessments were summarized. By using the Cochrane Collaboration’s tool, three RCTs revealed low risks of bias, and ten observational studies had high qualities (S1 Fig). Egger’s regression test did not reach significance for publication bias in any of the analysis (all $P$-values $> 0.05$) (S1 Table and S2–S4 Figs). Leave-one-out sensitivity analyses also showed

Table 1. (Continued)

| Studies       | Country | Sample size (Trt/ctrl) | Comparator antibiotic | Comorbidities of study population | Outcomes | Effect measures | Point estimates (95%CI) | Follow-up duration |
|---------------|---------|------------------------|-----------------------|----------------------------------|----------|----------------|------------------------|--------------------|
| Berni et al.  | UK      | 63,223/963,075         | AMX                   | General population               | ACM      | Rate ratio*   | 1.35 (1.22, 1.49)       | 37                 |
|               |         |                        |                       |                                  | AA       | Risk ratio    | 1.28 (1.12, 1.46)       | 37                 |

UK, United Kingdom; US, United States; ACM, all-cause mortality; CD, cardiac death; AMI, acute myocardial infarction; AA, arrhythmia alliance; Trt, treatment; Crtl, control; AMX, amoxicillin; CLR, clarithromycin; AMC, amoxicillin-Clavulanate; AZM, azithromycin; RCT, randomized controlled trial; AMI, acute myocardial infarction; COPD, chronic obstructive pulmonary disease; CAP, community acquired pneumonia

* Enrollees in all observational studies were indicated for either CLR or other antibiotics treatment due to infectious diseases (e.g. pneumonia and Helicobacter pylori infection).

§ Nested case-control study

Conversion from odds ratio (OR) to rate ratio (RR) via the equation of $RR = \frac{OR}{1 + OR \cdot r}$, $r$ is the outcome event rate for the comparator group

Conversion from hazard ratio (HR) to rate ratio (RR) via the equation of $RR = \frac{1}{e^{HR \cdot \ln(1/r)}}$, $r$ is the outcome event rate for the comparator group

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mortality (pooled RR = 1.19, 95% CI = 0.69–2.07) (Fig 3E), with a large heterogeneity ($I^2 = 87.9\%, P < 0.001$).
consistent and solid pooled estimates (S5 and S6 Figs). In subgroup analyses, the summary rate ratios of all-cause mortality were similar across different study types, study quality, and effect measures, although the heterogeneities were high in the subgroups of observational study \( (I^2 = 98.4\%) \) and non-HR group \( (I^2 = 97.8\%) \) (S2 Table).

### Discussion

In this meta-analysis, our results revealed no increased long-term all-cause mortality, comparing CLR and the other antibiotics, among patients with and without CAD. There was no increased risk of cardiac mortality, AMI, and arrhythmia, comparing CLR use versus alternative antibiotics use or placebo.

Our results suggested no significant increased long-term mortality by CLR use in concordance with the lack of possible mechanistic explanation of long-term effects of CLR.[25] The potential explanation for macrolides related arrhythmia is by inhibiting potassium current in cardiac muscle cells,[26] which very unlikely persist for years. Most human studies were consistent with the observation that no delayed effects result in ischemic heart diseases or cardiovascular death one year after treatment starts.[4, 21, 23]

However, previous meta-analysis by Cheng et al.[27] in 2015 concluded elevated risks of CV events and all-cause mortality among CLR users compared with non-macrolide antibiotics users. Similarly, in 2018 another meta-analysis summarized an elevated risk of MI comparing CLR versus non-macrolide antibiotics use.[3] Both meta-analyses did not include updated researches and had some methodological problems. First, comparing to our study, eight and six relevant articles were not included for analyses in those meta-analyses by Cheng et al. and Gorelik et al., respectively. Second, without considering the differences of follow-up duration and study designs, previous analyses failed to appropriately address time window of the adverse effect. In fact, they improperly combined the results from short-term and long-term follow-up duration.[27] Third, the use of odds ratio in the meta-analysis did not account for the time-to-event effects, which may substantially affect the estimates according to our results.

In order to properly interpret the summary risk of CLR, the study design (placebo versus alternative antibiotics controls), study populations (with versus without cardiovascular diseases), and follow-up durations (long-term versus short-term), which may potentially bias the results, should be discussed. The rationales behind the two RCT studies comparing CLR and placebo were not reasonable at present.[8, 18] In fact, a fallacy in the early 1990s discussed that Chlamydia pneumoniae was the pathogen leading to atherosclerosis, and treating this pathogen by CLR may reduce risks of CAD.[28] Thus, the RCT performed at that time compared all-cause and CV mortality in CLR users to placebo, instead of alternative antibiotics users. Nowadays, CLR is mainly used to treat infectious diseases rather than CAD and thus investigating...
cardiovascular risks among patients of infectious diseases by comparing CLR with alternative antibiotic treatment is more reasonable. [29]

Interpretation of meta-analysis that included in the CLARICOR trial should be cautious for the following reasons. [8] First, the two-week duration of CLR exposure was much shorter than...
the follow-up duration of 10 years. There is no plausible biological explanation for the 10-year effects by a short-term use of rapid acting antibiotic, particularly when no increased risk of mortality and AMI was found within 3 years of follow-up. This might suggest possible effects by other intermediates or factors after randomization. Second, only 32% of enrollees in the CLARICOR trial were randomized and there were imbalance factors (e.g., smoking) at baseline despite randomization, implying an invalid RCT design and potential biased results. Third, participants in the CLARICOR trial had no infectious disease indicated for CLR use, suggesting lack of generalizability to infection-diseased patients.

Our results from two cohort studies surprisingly showed that CLR users had a 63% increased rate ratio of all-cause mortality, but neither cardiac mortality nor arrhythmia, within 3 months of follow-up. The increased risk might result from non-cardiac mortality or potential bias by indication due to the cohort study design. By contrast, in 2017, one meta-analysis of cohort studies and RCTs revealed an increased risk of CV mortality, AMI, and arrhythmia among macrolides users after follow-up less than 30 days, but not in long-term follow-up. In our results of observational studies, there was no increased risk of all-cause mortality and AMI among CLR users with follow-up greater or equal to 1 year, suggesting the biological effects of CLR may diminish as time went by. Similarly, in CLARICOR trial, the measures of risk by CLR use were increased during 0–3 years of follow-up and returned toward the null during 6–10 years of follow-up. Given limited original studies investigating the short-term risks of CLR use so far, further clinical trials are warranted, particular in patients with cardio-vascular diseases.

There are a few potential limitations in this study. First, the number of studies that met our inclusion criteria was small. The diverse study designs and study populations make it hard to simply summarize the effect measures. However, our methods are much comprehensive compared to previous meta-analyses. Second, there were large heterogeneities in this meta-analysis (I² > 50%), possibly from various types of studies, selection of placebo or active comparators, and different baseline comorbidities. Thus, our subgroup findings provide more insights of risks of CLR users in different clinical settings.

Conclusions

Comparing CLR and alternative antibiotics use, overall, our results showed no increased long-term all-cause mortality among patients with and without CAD. However, CLR may be associated with increased risks of all-cause mortality in specific populations or within a short-term of follow-up period. Further RCTs to compare CLR versus alternative antibiotics use in a short-term period or among patients with CV diseases were highly suggested.

Supporting information

S1 File. The Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) checklist. (DOC)

S2 File. Supporting information of methods. Details in search keywords and quality assessment were provided. (DOCX)

S1 Table. Publication bias. The Egger’s tests showed no significant small study effects. (DOCX)
**S2 Table. Subgroup analysis.** The pooled rate ratios of all-cause mortality and heterogeneity were summarized by subgroups.

(DOCX)

**S1 Fig. Quality assessment.** The risks of bias for randomized controlled trials and observational studies were assessed by using the Cochrane Collaboration’s tool and the Newcastle-Ottawa Quality Assessment Scale (NOS), respectively.

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**S2 Fig.** (A) Egger’s test and (B) Funnel plot of all studies on all-cause mortality with long-term follow-up. The results showed no obvious publication bias.

(TIF)

**S3 Fig. Egger’s tests of the observational studies.** The results showed no obvious publication bias in studies with the outcomes of cardiac mortality after (A) short-term and (B) immediate follow-up durations and (C) studies with short-term outcome of arrhythmia.

(TIF)

**S4 Fig.** Egger’s tests of those observational studies with long-term outcomes: (A) all-cause mortality and (B) acute myocardial infarction. The results showed no obvious publication bias.

(TIF)

**S5 Fig. Leave-one-out analysis of randomized controlled trials and observational studies with long-term outcomes of all-cause mortality revealed no strong effects by any single study.**

(TIF)

**S6 Fig.** Leave-one-out analysis of observational studies with (A) short-term and (B) immediate outcomes of cardiac mortality showed no strong effects by any single study.

(TIF)

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Clarithromycin use and mortality risk

Writing – review & editing: Ching-Hui You, Cheng-Kuan Lin, Stefania I. Papatheodorou, Szu-Ta Chen.

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