A note on degenerate derangement polynomials and numbers

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Abstract: In this paper, we study the degenerate derangement polynomials and numbers, investigate some properties of those polynomials and numbers and explore their connections with the degenerate gamma distributions. In more detail, we derive their explicit expressions, recurrence relations and some identities involving the degenerate derangement polynomials and numbers and other special polynomials and numbers, which include the fully degenerate Bell polynomials, the degenerate Fubini polynomials and the degenerate Stirling numbers of both kinds. We also show that those polynomials and numbers are connected with the moments of some variants of the degenerate gamma distributions.

Keywords: degenerate derangement polynomials; degenerate gamma distribution; degenerate Fubini polynomials; fully degenerate Bell polynomials; degenerate Stirling numbers

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1. Introduction

A derangement is a permutation with no fixed points. In other words, a derangement is a permutation of the elements of a set that leaves no elements in their original places. The number of derangements of a set of size \( n \) is called the \( n \)-th derangement number and denoted by \( d_n \). The first few terms of the derangement number sequence \( \{d_n\}_n=0^\infty \) are \( d_0 = 1, \ d_1 = 0, \ d_2 = 1, \ d_3 = 2, \ d_4 = 9, \ldots \). It was Pierre Rémonde de Motmort who initiated the study of counting derangements in 1708 (see [1]).

Carlitz was the first one who studied degenerate versions of some special polynomials and numbers, namely the degenerate Bernoulli polynomials and numbers and degenerate Euler polynomials and numbers. In recent years, the study of various degenerate versions of some special polynomials and numbers regained the interests of quite a few mathematicians and yielded many interesting arithmetical
and combinatorial results. It is remarkable that the study of degenerate versions is not just limited to polynomials but can be extended to transcendental functions like gamma functions (see [9,14]).

The aim of this paper is to study the degenerate derangement polynomials, which are a degenerate version of the derangement polynomials. Here the derangement polynomials are a natural extension of the derangement numbers. In more detail, we derive their explicit expressions, recurrence relations and some identities involving those polynomials and numbers and other special polynomials and numbers, which include the fully degenerate Bell polynomials, the degenerate Fubini polynomials and the degenerate Stirling numbers of both kinds. We also introduce the higher-order degenerate derangement polynomials. Then we explore the degenerate gamma distributions as a degenerate version of the gamma distributions. We show that the moments of distributions coming from some variants of degenerate gamma distributions are related to the degenerate derangement polynomials or the degenerate derangement numbers or the higher-order degenerate derangement polynomials.

For the rest of this section, we recall the necessary facts about the degenerate derangement polynomials and numbers and the degenerate exponential functions.

As is well known, the generating function of the derangement numbers is given by

$$
\frac{1}{1-t} = \sum_{n=0}^{\infty} d_n t^n n!, \quad (\text{see } [1, 3, 4, 8, 12, 13]).
$$

From (1.1), we note that

$$
d_n = n! \sum_{i=0}^{n} \frac{(-1)^i}{i!}, \quad (n \geq 0), \quad (\text{see } [8, 10, 12, 13]).
$$

The derangement polynomials are defined by the generating function as

$$
\frac{1}{1-t} e^{(x-1)t} = \sum_{n=0}^{\infty} d_n(x) \frac{t^n}{n!}, \quad (\text{see } [12, 13]).
$$

By (1.3), we get

$$
d_n(x) = \sum_{l=0}^{n} \binom{n}{l} d_l x^{n-l},
$$

$$
= n! \sum_{l=0}^{n} \frac{(x-1)^l}{l!}, \quad (n \geq 0).
$$

Clearly, we have $d_n(0) = d_n$.

For any nonzero real number $\lambda$, the degenerate exponential function is defined as

$$
e^{\lambda t} = (1 + \lambda t)^{\lambda} = \sum_{n=0}^{\infty} \frac{(x)_{n,\lambda}}{n!} t^n, \quad (\text{see } [2, 5, 9, 11, 16]),
$$

where $(x)_{n,\lambda} = x(x-\lambda) \cdots (x-(n-1)\lambda), \quad (n \geq 1)$.

For brevity we denote $e^{\lambda t}$ by $e_{\lambda}(t)$. In this paper, we study the degenerate derangement polynomials which are derived from the degenerate exponential function.

From the definition of degenerate derangement polynomials, we investigate some properties and recurrence relations and new identities associated with special numbers and polynomials.
2. Degenerate derangement polynomials

In light of (1.3), we may consider the degenerate derangement polynomials which are given by

\[
\frac{1}{1 - te^{-x}} = \sum_{n=0}^{\infty} d_{n,\lambda}(x) \frac{t^n}{n!}.
\]  

(2.1)

When \( x = 0 \), \( d_{n,\lambda} = d_{n,\lambda}(0) \) are called the degenerate derangement numbers.

From (1.5) and (2.1), we get

\[
\sum_{n=0}^{\infty} d_{n,\lambda}(x) \frac{t^n}{n!} = \sum_{l=0}^{\infty} t^l \sum_{m=0}^{\infty} (x - 1)_{m,\lambda} \frac{t^m}{m!}
\]

\[
= \sum_{n=0}^{\infty} \left( n! \sum_{m=0}^{n} \frac{(x - 1)_{m,\lambda}}{m!} \right) \frac{t^n}{n!}.
\]  

(2.2)

Comparing the coefficients on both sides of (2.2), we obtain the following proposition.

**Proposition 1.** For \( n \geq 0 \), we have

\[
d_{n,\lambda}(x) = n! \sum_{l=0}^{n} \frac{(x - 1)_{l,\lambda}}{l!}.
\]

In particular, for \( x = 0 \), we obtain

\[
d_{n,\lambda} = n! \sum_{l=0}^{n} \frac{(-1)_{l,\lambda}}{l!}.
\]

Now, we observe that

\[
e^{-x}_{\lambda} = 1 + \sum_{n=1}^{\infty} \left( d_{n,\lambda}(x) - nd_{n-1,\lambda}(x) \right) \frac{t^n}{n!}.
\]  

(2.3)

From (1.5) and (2.3), we have

\[
(x - 1)_{n,\lambda} = d_{n,\lambda}(x) - nd_{n-1,\lambda}(x),
\]  

(2.4)

and

\[
(-1)_{n,\lambda} = d_{n,\lambda} - nd_{n-1,\lambda}, \quad (n \geq 1).
\]

In addition, by (2.1), we get

\[
d_{n,\lambda}(x) = \sum_{l=0}^{n} \binom{n}{l} d_{l,\lambda}(x)_{n-l,\lambda}, \quad (n \geq 0).
\]  

(2.5)

Therefore, by (2.5), we obtain the following theorem.

**Theorem 2.** The following identities hold true:

\[
d_{n,\lambda}(x) = \sum_{l=0}^{n} \binom{n}{l} d_{l,\lambda}(x)_{n-l,\lambda}, \quad (n \geq 0),
\]

\[
(x - 1)_{n,\lambda} = d_{n,\lambda}(x) - nd_{n-1,\lambda}(x), \quad (n \geq 1),
\]

\[
(-1)_{n,\lambda} = d_{n,\lambda} - nd_{n-1,\lambda}, \quad (n \geq 1).
\]
Replacing \( t \) by \( 1 - e_A(t) \) in (2.1), we get

\[
e^{x-1}(1 - e_A(t)) = e_A(t) \sum_{l=0}^{\infty} \frac{d_{l,A}(x)}{l!} (1 - e_A(t))^l = \sum_{m=0}^{\infty} \frac{(1)^{m,A}}{m!} \sum_{j=0}^{\infty} \frac{(-1)^j d_{j,A}(x) S_{2,A(j, l)} t^j}{j!}
\]

\[
= \sum_{n=0}^{\infty} \left( \sum_{j=0}^{\infty} \sum_{l=0}^{n} \binom{n}{j} (1)^{n-j,A} (-1)^j d_{j,A}(x) S_{2,A(j, l)} \right) \frac{t^n}{n!}.
\]

Here \( S_{2,A}(n, l), (n \geq l) \), are the degenerate Stirling numbers of the second kind given either by

\[
(x)_{n,A} = \sum_{l=0}^{n} S_{2,A}(n, l)(x), \quad (n \geq 0),
\]

or by

\[
\frac{1}{m!} (e_A(t) - 1)^m = \sum_{n=0}^{\infty} S_{2,A}(n, m) \frac{t^n}{n!}, \quad (m \geq 0), \quad (see [7]),
\]

where \((x)_0 = 1, (x)_n = x(x-1) \cdots (x-n+1), (n \geq 1)\). Alternatively, (2.6) is also given by

\[
e^{x-1}(1 - e_A(t)) = \sum_{m=0}^{\infty} \frac{(x-1)^m}{m!} (1 - e_A(t))^m = \sum_{n=0}^{\infty} \left( \sum_{m=0}^{n} \binom{n}{m} (x-1)^m (-1)^m S_{2,A}(n, m) \right) \frac{t^n}{n!}.
\]

Therefore, by (2.6) and (2.7), we obtain the following theorem.

**Theorem 3.** For \( n \geq 0 \), we have

\[
\sum_{j=0}^{n} \sum_{l=0}^{j} \binom{j}{l} (1)^{n-j,A} (-1)^j d_{l,A}(x) S_{2,A(j, l)} = \sum_{j=0}^{n} (x-1)^{n-j,A} (-1)^j S_{2,A}(n, j).
\]

Recently, the degenerate Fubini polynomials are introduced as

\[
\frac{1}{1 - y(e_A(t) - 1)} = \sum_{n=0}^{\infty} F_{n,A}(y) \frac{t^n}{n!}, \quad (see [11, 15]).
\]

(2.8)

Note that \( \lim_{A \to 0} F_{n,A}(y) = F_n(y) \) are the ordinary Fubini polynomials (see [6]). Replacing \( t \) by \( e_A(t) - 1 \) in (2.1), we get

\[
\frac{1}{2 - e_A(t)} e^{x-1}(e_A(t) - 1) = \sum_{l=0}^{\infty} \frac{d_{l,A}(x)}{l!} (e_A(t) - 1)^l
\]

\[
= \sum_{n=0}^{\infty} \left( \sum_{l=0}^{n} S_{2,A}(n, l)d_{l,A}(x) \right) \frac{t^n}{n!}.
\]

(2.9)
In terms of (2.8), we note that (2.9) is also given by

\[ \frac{1}{2 - e^t} e^{x-1}(e^t - 1) \]

(2.10)

\[ = \sum_{l=0}^{\infty} F_{l,\lambda}(1) \frac{t^l}{l!} \sum_{m=0}^{\infty} (x - 1)^m \frac{1}{m!} (e^t - 1)^m \]

\[ = \sum_{l=0}^{\infty} F_{l,\lambda}(1) \frac{t^l}{l!} \sum_{j=0}^{\infty} \sum_{m=0}^{\infty} (x - 1)^m S_{2,\lambda}(j, m) \frac{t^j}{j!} \]

\[ = \sum_{n=0}^{\infty} \sum_{l=0}^{n} \sum_{m=0}^{l} \binom{n}{l} F_{n-l,\lambda}(1)(x - 1)^m S_{2,\lambda}(l, m) \frac{t^n}{n!} \].

Therefore, by (2.9) and (2.10), we obtain the following theorem.

**Theorem 4.** For \( n \geq 0 \), we have

\[ \sum_{l=0}^{n} S_{2,\lambda}(n, l) d_{l,\lambda}(x) = \sum_{l=0}^{n} \sum_{m=0}^{l} \binom{n}{l} F_{n-l,\lambda}(1)(x - 1)^m S_{2,\lambda}(l, m). \]

Let \( \log_{\lambda}(t) \) be the compositional inverse function of \( e^t \). Recall that the degenerate Stirling numbers of the first kind are defined either by

\[ (x)_n = \sum_{l=0}^{n} S_{1,\lambda}(n, l)(x)_{l,\lambda}, \quad (n \geq 0), \]

or by

\[ \frac{1}{m!} (\log_{\lambda}(1 + t))^m = \sum_{n=m}^{\infty} S_{1,\lambda}(n, m) \frac{t^n}{n!}, \quad (m \geq 0), \quad (\text{see [7, 14]}). \]

Replacing \( t \) by \( \log_{\lambda}(1 + t) \) in (2.8) with \( y = 1 \), we get

\[ \frac{1}{1 - t} = \sum_{l=0}^{\infty} F_{l,\lambda}(1) \frac{1}{l!} (\log_{\lambda}(1 + t))^l \]

(2.11)

\[ = \sum_{n=0}^{\infty} \left( \sum_{l=0}^{n} F_{l,\lambda}(1) S_{1,\lambda}(n, l) \right) \frac{t^n}{n!}. \]

Writing the left hand side of (2.11) differently, we have

\[ \frac{1}{1 - t} = \left( \frac{1}{1 - t} e^{t-1}(t) \right) e^t \]

(2.12)

\[ = \sum_{l=0}^{\infty} d_{l,\lambda} \frac{t^l}{l!} \sum_{m=0}^{\infty} \frac{(1)_{m,\lambda}}{m!} \frac{t^m}{m!} \]

\[ = \sum_{n=0}^{\infty} \left( \sum_{l=0}^{n} \binom{n}{l} d_{l,\lambda}(1) \frac{t^n}{n!} \right). \]

Therefore, by (2.1), (2.11) and (2.12), we obtain the following theorem.
Theorem 5. For \( n \geq 0 \), we have

\[
 n! = \sum_{l=0}^{n} F_{l,\lambda}(1) S_{1,\lambda}(n, l) = \sum_{l=0}^{n} \binom{n}{l} d_{l,\lambda}(1)_{n-l,\lambda} = \sum_{l=0}^{n} \binom{n}{l} d_{l,\lambda}(x)(1 - x)_{n-l,\lambda}.
\]

Recently, Kim-Kim considered the fully degenerate Bell polynomials given by

\[
e_{\lambda}(x(e_{\lambda}(t) - 1)) = \sum_{n=0}^{\infty} \text{Bel}_{n,\lambda}(x) \frac{t^n}{n!}, \quad \text{(see [2])},
\]

(2.13)

Replacing \( t \) by \( \log_{\lambda}(1 + t) \) in (2.13) with \( x = 1 \), we get

\[
e_{\lambda}(t) = \sum_{m=0}^{\infty} \text{Bel}_{m,\lambda} \frac{1}{m!} (\log_{\lambda}(1 + t))^m \quad \text{(2.14)}
\]

\[
e_{\lambda}(t) = \sum_{m=0}^{\infty} \text{Bel}_{m,\lambda} \sum_{n=m}^{\infty} S_{1,\lambda}(n, m) \frac{t^n}{n!} = \sum_{n=0}^{\infty} \left( \sum_{m=0}^{n} \text{Bel}_{m,\lambda} S_{1,\lambda}(n, m) \right) \frac{t^n}{n!}.
\]

Obviously, (2.14) is also given by

\[
e_{\lambda}(t) = \sum_{n=0}^{\infty} (1)_{n,\lambda} \frac{t^n}{n!}.
\]

(2.15)

Therefore, by (2.14) and (2.15), we obtain the following theorem.

Theorem 6. For \( n \geq 0 \), we have

\[
(1)_{n,\lambda} = \sum_{m=0}^{n} \text{Bel}_{m,\lambda} S_{1,\lambda}(n, m),
\]

and

\[
\text{Bel}_{n,\lambda} = \sum_{m=0}^{n} (1)_{m,\lambda} S_{2,\lambda}(n, m).
\]

We observe that

\[
\frac{1}{1 - t} = e_{\lambda}^{-1}(\log_{\lambda}(1 - t)) = \sum_{m=0}^{\infty} (-1)_{m,\lambda} \frac{1}{m!} (\log_{\lambda}(1 - t))^m = \sum_{m=0}^{\infty} (1)_{m,\lambda} \left( \sum_{n=m}^{\infty} (-1)^n S_{1,\lambda}(n, m) \frac{t^n}{n!} \right) \frac{1}{n!}.
\]

(2.16)

From Theorem 5 and (2.16), we obtain

\[
n! = (-1)^n \sum_{m=0}^{n} (-1)_{m,\lambda} S_{1,\lambda}(n, m) = \sum_{m=0}^{n} \binom{n}{m} d_{m,\lambda}(x)(1 - x)_{n-m,\lambda}.
\]
Replacing \( t \) by \( \log_\lambda(1 - t) \) in (2.13) with \( x = 1 \), we get

\[
e^{-\lambda}(t) = \sum_{k=0}^{\infty} \frac{1}{k!} (\log_\lambda(1 - t))^k
\]

\[
= \sum_{k=0}^{\infty} \sum_{n=k}^{\infty} S_{1,\lambda}(n, k) (-1)^n \frac{t^n}{n!}
\]

\[
= \sum_{n=0}^{\infty} \left( \sum_{k=0}^{n} \frac{1}{k!} \right) S_{1,\lambda}(n, k) (-1)^n \frac{t^n}{n!}.
\]

(2.17)

We remark that (2.17) is alternatively given by

\[
e^{-\lambda}(t) = e^{-1} - \lambda t = \sum_{n=0}^{\infty} \frac{(-1)^n}{n!} \lambda^n t^n.
\]

(2.18)

Thus, from (2.17) and (2.18), we have

\[
\sum_{k=0}^{n} \frac{1}{k!} \right) S_{1,\lambda}(n, k) = (-1)^n (-1)^n_{n, -\lambda}, \quad (n \geq 0).
\]

(2.19)

Replacing \( t \) by \( 1 - e^{-\lambda}(t) \) in (2.1) with \( x = 0 \), we get

\[
e^{-\lambda}(t) = e^{-\lambda}(t) = \sum_{n=0}^{\infty} \frac{1}{n!} (-1)^n (e^{-\lambda}(t) - 1)^n
\]

(2.20)

An alternative expression of (2.20) is given by

\[
e^{-\lambda}(t) = e^{-\lambda}(t) - \lambda t = \sum_{n=0}^{\infty} \frac{1}{n!} (-1)^n \lambda^n t^n.
\]

(2.21)

From (2.20) and (2.21), we have

\[
\sum_{n=0}^{\infty} (-1)^n d_{m,\lambda} S_{2,\lambda}(n, m) t^n
\]

(2.22)

Therefore, by (2.19) and (2.22), we obtain the following theorem.
Theorem 7. For $n \geq 0$, we have
\[
\sum_{m=0}^{n} (-1)^m d_{m,\lambda} S_{2,\lambda}(n, m) = \sum_{m=0}^{n} \binom{n}{m} \text{Bel}_{m,\lambda} (-1)_{n-m,\lambda}.
\]

In addition, we have
\[
\sum_{k=0}^{n} \text{Bel}_{k,\lambda} S_{1,\lambda}(n, k) = (-1)^n (-1)_{n,\lambda}, \quad (n \geq 0).
\]

For $r \in \mathbb{N}$, we define the degenerate derangement polynomials of order $r$ which are given by
\[
\frac{1}{(1-t)^r} e^{x-1}(t) = \sum_{n=0}^{\infty} d_n^{(r)}(x) \frac{t^n}{n!}.
\] (2.23)

When $x = 0$, $d_n^{(r)}(0)$ are called the degenerate derangement numbers of order $r$.

From (2.23), we note that
\[
\sum_{n=0}^{\infty} d_n^{(r)}(x) \frac{t^n}{n!} = \sum_{m=0}^{\infty} \binom{r+m-1}{m} \sum_{l=0}^{\infty} (x-1)_{l,\lambda} \frac{t^l}{l!}.
\] (2.24)

Comparing the coefficients on both sides of (2.24), we obtain the following theorem.

Theorem 8. For $n \geq 0$, we have
\[
d_n^{(r)}(x) = n! \sum_{l=0}^{n} \frac{(x-1)_{l,\lambda}}{l!} \left( r + n - l - 1 \atop n - l \right).
\]

In particular, for $x = 0$, we have
\[
d_n^{(r)} = n! \sum_{l=0}^{n} (-1)_{l,\lambda} \left( r + n - l - 1 \atop n - l \right).
\]

By (2.1), we get
\[
\frac{1}{1+t} e_{\lambda}^{-1}(-t) = \sum_{m=0}^{\infty} d_{m,\lambda} (-1)^m \frac{t^m}{m!}.
\] (2.25)

Replacing $t$ by $e_{\lambda}(t) - 1$ in (2.25), we get
\[
e_{\lambda}^{-1}(1 - e_{\lambda}(t)) = e_{\lambda}(t) \sum_{m=0}^{\infty} d_{m,\lambda} (-1)^m \frac{1}{m!} (e_{\lambda}(t) - 1)^m
\] (2.26)
\[
e_{\lambda}(t) \sum_{m=0}^{\infty} d_{m,\lambda} (-1)^m \sum_{j=m}^{\infty} S_{2,\lambda}(j, m) \frac{t^j}{j!}
\]
\[
\sum_{j=0}^{\infty} \frac{j^n}{j!} \sum_{m=0}^{\infty} (-1)^m d_{m,A} S_{2,-A}(j, m) \frac{t^j}{j!}
= \sum_{j=0}^{\infty} \sum_{m=0}^{n} \binom{n}{j} (1)_{n-j,-A} (-1)^m d_{m,A} S_{2,-A}(j, m) \frac{t^j}{j!}.
\]

Alternatively, (2.26) is also given by

\[
e^{-1} (1 - e_{-A}(t)) = e_{-A}(e_{-A}(t) - 1) = \sum_{n=0}^{\infty} \text{Bel}_{n,-A}(1) \frac{t^n}{n!}.
\]

Therefore, by (2.26) and (2.27), we obtain the following theorem.

**Theorem 9.** For \( n \geq 0 \), we have

\[
\text{Bel}_{n,-A}(1) = \sum_{j=0}^{n} \sum_{m=0}^{j} \binom{n}{j} (1)_{n-j,-A} (-1)^m d_{m,A} S_{2,-A}(j, m).
\]

### 3. Further remarks

Let \( f(x) \) be the probability density function of the continuous random variable \( X \), and let \( g(x) \) be a real valued function. Then the expectation of \( g(X) \), \( E[g(X)] \), is defined by

\[
E[g(X)] = \int_{-\infty}^{\infty} g(x)f(x)dx, \quad \text{(see [18])}.
\]

A continuous random variable \( X \), whose density function is given by

\[
f(x) = \begin{cases} 
\beta e^{-\beta x} \frac{(\beta x)^{\alpha-1}}{\Gamma(\alpha)}, & \text{if } x \geq 0, \\
0, & \text{if } x < 0,
\end{cases}
\]

for some \( \beta > 0 \) and \( \alpha > 0 \), is said to be the gamma random variable with parameters \( \alpha, \beta \) and denoted by \( X \sim \Gamma(\alpha, \beta) \).

Let \( X \sim \Gamma(1, 1) \). Then, for all \( t < 1 \), we have

\[
E[e^{xt} \cdot e_{-A}(t)] = e_{-A}^{-1}(t) \int_{0}^{\infty} e^{xt} e^{-A} dx
= \frac{1}{1 - t} e_{-A}^{-1}(t) = \sum_{n=0}^{\infty} d_{n,A} t^n.
\]

Clearly, we also have

\[
E[e^{xt} e_{-A}(t)] = \sum_{n=0}^{\infty} \left( \sum_{m=0}^{n} \binom{n}{m} (-1)^{n-m} E[X^m] \right) \frac{t^n}{n!}.
\]

Therefore, by (3.3) and (3.4), we obtain the following equations.
For \( n \geq 0 \), we have
\[
\sum_{m=0}^{n} \binom{n}{m} (-1)^{n-m} E[X^m] = d_{n,\lambda},
\]
and, more generally, we also have
\[
\sum_{m=0}^{n} \binom{n}{m} (x-1)^{n-m} E[X^m] = d_{n,\lambda}(x).
\]

Unless otherwise stated, for the rest of this section we assume that \( \lambda \in (0, 1) \). The degenerate gamma function \( \Gamma_{\lambda}(x) \), which is initially defined for \( 0 < \text{Re}(s) < \frac{1}{\lambda} \) by the following integral
\[
\Gamma_{\lambda}(s) = \int_{0}^{\infty} e^{-t} t^{s-1} dt, \quad \text{(see [9, 14]), (3.5)}
\]
can be continued to a meromorphic function on \( \mathbb{C} \), whose only singularities are simple poles at \( s = 0, -1, -2, \ldots, \frac{1}{\lambda}, \frac{1}{\lambda}+1, \frac{1}{\lambda}+2, \ldots \). Thus, by (3.5), we get
\[
\Gamma_{\lambda}(k) = \frac{\Gamma(k)}{(1)_{k+1,\lambda}}, \quad \left( k \in \mathbb{N}, \lambda \in (0, \frac{1}{k}) \right),
\]
and, in particular, we have
\[
\Gamma_{\lambda}(1) = \frac{1}{1 - \lambda}, \quad \text{(see [9]).}
\]

A random variable \( X = X_{\lambda} \) is said to have the degenerate gamma distribution with parameters \( \alpha \) and \( \beta \), \( (\frac{1}{\lambda} > \alpha > 0, \beta > 0) \), and denoted by \( X \sim \Gamma_{\lambda}(\alpha, \beta) \), if its probability density function has the form
\[
f_{\lambda}(x) = \begin{cases} \frac{1}{\Gamma(\alpha)\beta^{\alpha-1}} e^{-\lambda} e^{\beta x}, & \text{if } x \geq 0, \\ 0, & \text{otherwise.} \end{cases}
\]

Note that \( \frac{d}{dx} e_{\lambda}^{\beta}(x) = ce_{\lambda}^{\beta-\lambda}(x) \), for any constant \( c \). Then, for \( X \sim \Gamma_{\lambda}(1, 1) \), we have
\[
E[e^{-\lambda}(X)] = (1 - \lambda) \int_{0}^{\infty} e^{-\lambda}(x) e^{-1}(x) \, dx 
\]
\[
= (1 - \lambda) \int_{0}^{\infty} e^{-1-\lambda}(x) \, dx = \frac{1}{1 - \lambda} \frac{1}{1 - t} e_{\lambda}^{-1}(t) e_{\lambda}(t) 
\]
\[
= (1 - \lambda) \sum_{l=0}^{\infty} \frac{d_{l+1,\lambda}}{l!} \sum_{m=0}^{\infty} (1)_{m,\lambda} \frac{r^{m}}{m!} 
\]
\[
= \sum_{n=0}^{\infty} (1 - \lambda) \sum_{l=0}^{n} \frac{d_{l+1,\lambda}(1)_{n-l}}{l!} \frac{r^{n}}{n!}.
\]

Evidently, we also have
\[
E[e^{-\lambda}(X)] = E\left[ \frac{1}{1 + \lambda X} (1 + \lambda X)^{\frac{r}{n}} \right] 
\]
\[
= \sum_{n=0}^{\infty} E\left[ \frac{1}{1 + \lambda X} \left( \frac{1}{\lambda} \log(1 + \lambda X) \right)^{n} \right] \frac{r^{n}}{n!}.
\]

Therefore, (3.7) and (3.8), we obtain the following theorem.
Theorem 10. For $X \sim \Gamma_1(1, 1)$, we have

$$E\left[ \frac{1}{1 + \lambda X} \left( \frac{1}{\lambda} \log(1 + \lambda X) \right)^n \right] = (1 - \lambda) \sum_{l=0}^{n} d_{l,A}(1)_{n-l,A} \binom{n}{l}.$$

Now, we observe that

$$(\log(1 + \lambda X))^n = n! \sum_{m=n}^{\infty} S_1(m, n) \frac{\lambda^m}{m!} X^m, \quad (n \geq 0),$$

where $S_1(n, m)$ are the Stirling numbers of the first kind, (see [17,19,20]). In turn, we have

$$E\left[ \frac{1}{1 + \lambda X} \left( \frac{1}{\lambda} \log(1 + \lambda X) \right)^n \right] = \frac{n!}{\lambda^n} \sum_{m=n}^{\infty} S_1(m, n) \frac{\lambda^m}{m!} E\left[ \frac{X^m}{1 + \lambda X} \right]. \quad (3.9)$$

From Theorem 11 and (3.9), we have

$$\sum_{n=m}^{\infty} S_1(n, m) \frac{\lambda^m}{m!} E\left[ \frac{X^m}{1 + \lambda X} \right] = (1 - \lambda) \frac{n!}{\lambda^n} \sum_{l=0}^{n} d_{l,A}(1)_{n-l,A} \binom{n}{l}, \quad (n \geq 0),$$

where $X \sim \Gamma_1(1, 1)$.

For $X_1, X_2, \ldots, X_r \sim \Gamma(1, 1)$, assume that $X_1, X_2, \ldots, X_r$ are independent. Then we have

$$E\left[ e^{(X_1 + X_2 + \cdots + X_r) t} \right] = E\left[ e^{X_1 t} \right] E\left[ e^{X_2 t} \right] \cdots E\left[ e^{X_r t} \right] \cdot e^{(x-1) t} \left( \frac{1}{1 - t} \right)^r \times \left( \frac{1}{1 - t} \right) \times \cdots \times \left( \frac{1}{1 - t} \right)$$

$$= \sum_{n=0}^{\infty} d_r^{(x)}(x) \frac{t^n}{n!}. \quad (3.10)$$

Alternatively, (3.10) is given by

$$E\left[ e^{(X_1 + X_2 + \cdots + X_r) t} \right] = \sum_{l=0}^{\infty} E\left[ (X_1 + \cdots + X_r)^l \right] \frac{t^l}{l!} \sum_{m=0}^{\infty} (x - 1)_{m,A} \frac{t^m}{m!}$$

$$= \sum_{n=0}^{\infty} \left( \sum_{l=0}^{n} \binom{n}{l} E\left[ (X_1 + \cdots + X_r)^l \right] (x - 1)_{n-l,A} \right) \frac{t^n}{n!}. \quad (3.11)$$

By (3.10) and (3.11), we get

$$d_{n,A}(x) = \sum_{l=0}^{n} \binom{n}{l} E\left[ (X_1 + \cdots + X_r)^l \right] (x - 1)_{n-l,A}, \quad (n \geq 0).$$

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4. Conclusions

In this paper, we have dealt with the degenerate derangement polynomials \( d_{n,\lambda}(x) \), which are a degenerate version of the derangement polynomials \( d_n(x) \). We derived their explicit expressions, recurrence relations and some identities involving those polynomials and numbers and other special polynomials and numbers such as the fully degenerate Bell polynomials, the degenerate Fubini polynomials and the degenerate Stirling numbers of both kinds. We also introduced the higher-order degenerate derangement polynomials. Then we explored the degenerate gamma distributions as a degenerate version of the gamma distributions and showed that the moments of distributions coming from some variants of degenerate gamma distributions are related to the degenerate derangement polynomials or the degenerate derangement numbers or the higher-order degenerate derangement polynomials.

In recent years, the study of many special numbers and polynomials has been carried out by using several different methods, which include generating functions, combinatorial methods, umbral calculus, \( p \)-adic analysis, probability theory, special functions and differential equations. Moreover, the same has been done for various degenerate versions of quite a few special numbers and polynomials. Motivations for studying degenerate versions arise from their interests not only in combinatorial and arithmetical properties but also in their applications to symmetric identities, differential equations and probability theories.

It is one of our future projects to continue to investigate many ordinary and degenerate special numbers and polynomials by various means and to find their applications in physics, science and engineering as well as in mathematics.

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Conflict of interest

The authors declare no conflict of interest.

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