Coevolution of metallicity and star formation in galaxies to $z \simeq 3.7$: I. A Fundamental Plane

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draft version 10 May 2016

ABSTRACT

With the aim of understanding the coevolution of star formation rate (SFR), stellar mass ($M_\ast$), and oxygen abundance ($O/H$) in galaxies up to redshift $z \simeq 3.7$, we have compiled the largest available dataset for studying Metallicity Evolution and Galaxy Assembly (MEGA); it comprises $\sim 1000$ galaxies with a common $O/H$ calibration and spans almost two orders of magnitude in metallicity, a factor of $\sim 10^6$ in SFR, and a factor of $\sim 10^5$ in stellar mass. From a Principal Component Analysis, we find that the 3-dimensional parameter space reduces to a Fundamental Plane of Metallicity (FPZ) given by $12 + \log(O/H) = -0.14 \log(SFR) + 0.37 \log(M_\ast) + 4.82$. The mean $O/H$ FPZ residuals are small (0.16 dex) and consistent with trends found in smaller galaxy samples with more limited ranges in $M_\ast$, SFR, and $O/H$. Importantly, the FPZ is found to be approximately redshift-invariant within the uncertainties. In a companion paper, these results are interpreted with an updated version of the model presented by Dayal et al. (2013).

Key words: galaxies: evolution – galaxies: abundances – galaxies: star formation – galaxies: high redshift

1 INTRODUCTION

Galaxies are assembled over cosmic time by the accumulation of stellar mass ($M_\ast$) through star-formation (SF) processes. This build-up is accompanied by an increase of metal content, typically measured through the gas-phase oxygen abundance ($O/H$), the most abundant heavy element produced by massive stars. Stellar mass is a measure of the integrated SF activity over the history of the galaxy, while the star-formation rate (SFR) indicates the current rate for conversion of gas into stars. The gas-phase metallicity ($Z$) reflects not only the metal production from high-mass stars, but also the level of galaxy interactions with environment through inflows and outflows in the form of galactic winds.

Given the causal relation between star-formation processes and metal content in galaxies, it is not surprising that $M_\ast$, SFR, and $O/H$ are mutually correlated. The mass-metallicity relation (MZR, e.g., Tremonti et al. 2004) is a manifestation of the $M_\ast$–$Z$ correlation; the SF “main sequence” relates $M_\ast$ and SFR (SFMS, e.g., Brinchmann et al. 2004; Salim et al. 2007; Noeske et al. 2007). The mutual relations among the three variables extend to specific SFR ($sSFR \equiv SFR/M_\ast$) and metallicity which are also correlated (e.g., Salim et al. 2014; Yates & Kauffmann 2014).

These mutual correlations imply that residuals from the main relations (MZR, SFMS) should be correlated with the third variable. Indeed, from an analysis of data from the Sloan Digital Sky Survey (SDSS), Mannucci et al. (2010) found an expression that connected the residuals in the MZR to SFR; this was dubbed the “Fundamental Metallicity Relation” (FMR) and reduced the scatter in $O/H$ over $\sim 80000$ galaxies from $\sim0.1$ dex to 0.05–0.06 dex. In a similar vein, Lara-López et al. (2010) showed that the 3D space of $M_\ast$, SFR, and $O/H$ for $\sim 33000$ SDSS galaxies could be expressed through a two-dimensional (planar) surface (“Fundamental Plane”, FP). By fitting regressions to parameter pairs, they expressed the FP in terms of $M_\ast$ and found a residual variation of $\sim 0.16$ dex, larger however than that found for the FMR.

Given that reducing a three-dimensional (3D) parameter space to a (2D) plane is mathematically equivalent to diagonalizing the 3D covariance matrix, a natural approach to this problem is a Principal Component Analysis (PCA). A PCA was first applied to $M_\ast$, SFR, and $O/H$ by Hunt et al. (2012) for $\sim 1000$ galaxies from $z \sim 0 - 3.5$ selected to span a range of $\geq 10^5$ in SFR and two orders of magnitude
in O/H\(^1\). The PCA showed that the principal component dominated by O/H was the component most dependent on the other two, as by itself it comprised only \(\sim\)%2 of the total variance. The PCA resulted in a FP in metallicity (FPZ) with a spread of 0.17 dex in O/H, despite the vast range in the original parameters, including redshift. This FPZ applied to the same SDSS samples used by Mannucci et al. (2010, here mass limited) gave roughly the same residuals as the FMR, 0.06 dex. Thus, Hunt et al. (2012) concluded that the FPZ could be used to estimate metallicities with an accuracy of \(\sim 40–50\%\) over an extended range of M, and SFR, and moreover was a good representation of O/H at \(\sim 3\). It is now well established that both the MZR and the SFMS extend to the highest redshifts examined so far, but with differing normalizations relative to the Local Universe; at a given M, SFR (and sSFR) increases with increasing redshift (e.g., Noeske et al. 2007; Elbaz et al. 2011; Karim et al. 2011; Wuyts et al. 2011; Speagle et al. 2014) while metallicity decreases (e.g., Erb et al. 2006a; Maiolino et al. 2008; Mannucci et al. 2009; Cresci et al. 2012; Yabe et al. 2012; Henry et al. 2013; Cullen et al. 2014; Zahid et al. 2014; Troncoso et al. 2014; Steidel et al. 2014; Wuyts et al. 2014; Ly et al. 2015; de los Reyes et al. 2015). Consequently, if we assume that the FPZ is redshift-invariant (an assumption that we shall reassess below), the higher sSFRs found in high-z galaxy populations must be related, perhaps causally, to the lower metallicities observed at the same redshift. This is the hypothesis we examine in this paper.

In order to observationally constrain the evolution of metallicity with redshift, we have compiled a new dataset of \(\sim 1000\) star-forming galaxies from \(z \simeq 0\) to \(z \sim 3.7\) with nebular oxygen abundance measurements; we will refer this compilation as the “MEGA” dataset, corresponding to Metallicity Evolution and Galaxy Assembly. This compilation is a radical improvement over the dataset used by Hunt et al. (2012) because of the inclusion of several more high-z samples and, more importantly, because of a common metallicity calibration. Section 2 describes the 19 individual samples which form the MEGA compilation, together with our estimates of stellar masses and SFRs for the samples at \(z \simeq 0\). The procedures for aligning the individual samples to a common O/H calibration are outlined in Sect. 3. Sect. 4 describes the scaling relations for the MEGA dataset and re-evaluates the redshift invariance of the FPZ through a linear analysis of the correlations of M, SFR, and O/H in the MEGA sample and in \(\sim 80000\) galaxies at \(z \sim 0\) selected from the SDSS by Mannucci et al. (2010). The coevolution of SFR and O/H with redshift in the MEGA dataset is presented in Section 5, together with a comparison of results with previous work. We discuss our results and summarize our conclusions in Sect. 6. Throughout the paper we use a Chabrier (2003) Initial-Mass Function (IMF) and, when necessary, adopt the conversions for M, and SFR given by Speagle et al. (2014).

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1 To avoid problems with the curvature of the MZR at high metallicities, M was limited to \(< 10.5 \times M_\odot\) so those results are formally applicable only to galaxies less massive than this limit.

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2 GALAXY SAMPLES

Because of the need to compare stellar mass, M, SFR, and metal abundance [as defined by the nebular oxygen abundance, 12+log(O/H)], we have selected only samples of galaxies for which either these quantities are already available in the literature, or can be derived from published data. Here we discuss the estimates of M and SFR; the metallicity determinations for the samples will be discussed in Sect. 3.

2.1 Local Universe

Four samples of galaxies in the Local Universe met these criteria: the 11 Mpc distance-limited sample of nearby galaxies or Local Volume Legacy (11HUGS, LVL: Kennicutt et al. 2008; Lee et al. 2009, 2011); the Key Insights into Nearby Galaxies: a Far-Infrared Survey with Herschel (KINGFISH, Kennicutt et al. 2011); the starburst sample studied by Engelbracht et al. (2008), and the blue compact dwarf (BCD) sample by Hunt et al. (2010). There are 15 galaxies that appear both in the KINGFISH and LVL samples; for these, we used the KINGFISH parameters from Kennicutt et al. (2011) because of the uniform O/H calibration given by Moustakas et al. (2010). The starbursts from Engelbracht et al. (2008) were restricted to only those galaxies (42) with metallicities derived from the “direct” or “T,” method based on electron temperatures (see Table 1 and Sect. 3); the BCDs (23) all have \(T_\mathrm{e}\)-measured metallicities.

2.1.1 Star-formation rates

In order to maximize consistency, we have recalculated SFRs and M, for the four local samples starting from photometric fluxes reported in the literature. SFRs were derived according to Murphy et al. (2011) using for KINGFISH and LVL the hybrid method with far-ultraviolet (FUV)+total infrared luminosity (\(L_{\text{TIR}}\)); these data were available for 123 (of 138 non-KINGFISH) galaxies with O/H in the LVL and for 50 (of 55) KINGFISH galaxies. For the KINGFISH and LVL galaxies without these data, we adopted other SFR calibrations given by Murphy et al. (2011) including TIR (5 galaxies in KINGFISH, 6 LVL), UV (3 LVL), and Ho+24 \(\mu\)m (4 LVL). \(L_{\text{TIR}}\) was calculated according to Draine & Li (2007) and fluxes were taken from Dale et al. (2007, 2009) and Lee et al. (2009, 2011). For 20 LVL galaxies, the SFRs inferred from \(L_{\text{TIR}}\) were larger than those from FUV+\(L_{\text{TIR}}\) using the prescriptions by Murphy et al. (2011); in those cases, we adopted SFR(\(L_{\text{TIR}}\)). For NGC 253 and M 82, SFR(\(L_{\text{TIR}}\)) is \(\sim 2\) times SFR(FUV+\(L_{\text{TIR}}\)), but for the other galaxies the two estimates agree to within 30%. We also compared for the LVL galaxies the SFRs calculated with FUV+\(L_{\text{TIR}}\) with those inferred by combining Ho and 24-\(\mu\)m luminosities (\(L_{\text{H}\alpha}, L_{\text{24}}\); Calzetti et al. 2010; Murphy et al. 2011); SFR(FUV+\(L_{\text{TIR}}\)) tends to be \(\sim 1.6\) times larger than SFR(Ho+\(L_{\text{24}}\)), with a scatter of \(\sim 0.2\) dex. This is consistent with the findings of Leroy et al. (2012) who found a similar trend at low surface SFR densities such as those in the LVL galaxies.

Because FUV data are generally not available for the starbursts or the BCDs, for these we adopted the hybrid combination of Ho+\(L_{\text{24}}\) as prescribed by Murphy et
al. (2011). Total Ho fluxes were taken from Dopita et al. (2002); Gil de Paz et al. (2003); James et al. (2004); Pustilnik et al. (2004); Cannon et al. (2005); Moustakas & Kennicutt (2006); Schmitt et al. (2006); Lópe–Sánchez & Esteban (2008); Kennicutt et al. (2008); Cairós et al. (2010); James et al. (2010) and 24 μm measurements from Engelbracht et al. (2008). When these data were unavailable (9 galaxies), we adopted the SFR(TIR) prescription by Murphy et al. (2011) using the fluxes by Engelbracht et al. (2008). For SBS0335−052 and IzW18, we adopted the SFRs from radio-free-free emission (Hunt et al. 2004, 2005; Johnson et al. 2009), given the superiority of such estimates over other methods (e.g., Murphy et al. 2011). For one galaxy in the starburst sample (UM420), because of the lack of MIPS observations, the SFR was estimated from Ho luminosities, and for one galaxy (ESO 489−G56) there were no data available from which to infer SFR (so it was not considered further). For the (23) BCDs from Hunt et al. (2010), total Ho fluxes were taken from Gil de Paz et al. (2003); Rosagonzález et al. (2007); Pérez-Montero et al. (2011); Lagos et al. (2014), and 24 μm fluxes from Hunt et al. (2016, in prep.). As for the starbursts, for the three galaxies without Ho data, we used SFR(TIR); and for SBS1030+583 there were no MIPS data so we adopted SFR(Hα).

Considering the different SFR estimators discussed above, and considering their varying degrees of applicability, for the local samples the uncertainties on the SFRs are probably around a factor of 2 (0.3 dex). As mentioned above, the SFRs have been reported to a Chabrier (2003) IMF.

2.1.2 Stellar masses

We calculated the stellar masses according to Wen et al. (2013), a method based on WISE W1 (3.4 μm) luminosities. This approach exploits the approximately constant mass-to-light ratios of stellar populations at near-infrared wavelengths, independently of metallicity and age (Norris et al. 2014; McGaugh & Schombert 2014). However, when W1 photometry was not available, we used IRAC 3.6 μm photometry instead. In fact, the two bands are very similar; using data from Brown et al. (2014), Grossi et al. (2015) find for spirals a mean flux ratio $F_{3.6}/F_{3.4} = 1.02 ± 0.035$. Including also the data for dwarf irregulars from Brown et al. (2014) we find a mean flux ratio $F_{3.4}/F_{3.6} = 0.98±0.061$. Thus, we conclude that the ratio of the W1 and IRAC 3.6 μm bands is unity, with 5-6% scatter for galaxies like our targets.

For the starburst and BCD samples, we used the Higalactic formulation by Wen et al. (2013), rather than what they found for their full sample; the Higalactic galaxies have the lowest mass-to-light ratios in their compilation, corresponding roughly to the bluest regions of the galaxies studied by Zibetti et al. (2009). To better take into account the weak trends with abundance found by Wen et al. (2013), we also applied an approximate correction for low metallicity (by multiplying the mass-to-light ratio by 0.8 when $12+\log(O/H)\leq8.2$: Wen et al. 2013, see their Fig. 17). Instead, for the LVL and KINGFISH samples, we adopted the Wen et al. (2013) formulation based on morphological type, and considered an “early-type” galaxy one with Hubble type $T < 2^2$.

However, before applying the relations by Wen et al. (2013), we first subtracted nebular emission and emission from hot dust where possible. In starbursts and BCDs, such contamination can be very important in the near-infrared and can contribute 50% or more to the observed flux at these wavelengths (Hunt et al. 2001, 2002; Smith & Hancock 2009; Hunt et al. 2012). The ionized gas continuum contribution to the 3.4-3.6 μm flux was estimated from the SFR using the emission coefficients from Osterbrock & Ferland (2006). When possible, we also subtracted the hot-dust component, with the assumption that H-band emission is entirely stellar. Because H-band photometry is available for some of our sample, we used the data from Brown et al. (2014) to estimate the maximum possible IRAC 3.6 μm/H-band ratio in galaxies similar to our targets; 95% of the spiral/dwarf irregular galaxies have a flux ratio <2.4. This corresponds to a (Vega-based) [H-3.6] color of ~0.8, consistent with what is found for the pure stellar component in star-forming galaxies (Hunt et al. 2002). After subtraction of the nebular component, any excess over this ratio was attributed to hot dust and subtracted; this subtraction was not possible for 33 galaxies, including all the BCDs.

We compared the stellar masses obtained with the formulation of Wen et al. (2013) to those calculated according to Lee et al. (2006) based on IRAC 4.5 μm luminosities (used by Hunt et al. 2012). For the BCDs, the masses based on Wen et al. (2013) are on average ~0.2 dex lower than those based on Lee et al. (2006) with a scatter of 0.15 dex; this is not unexpected given the blue colors of these galaxies and the results of Zibetti et al. (2009) who showed that the Bell & de Jong (2001) calibration used by Lee et al. (2006) gives mass-to-light ratios that are too high for such blue galaxies. Instead for the starbursts the two estimates are in closer agreement, with ~0.06 dex difference on average and a scatter of 0.20 dex. Moreover, for the galaxies having both W1 and IRAC data, the stellar masses obtained from 3.6 μm luminosities are within ~5% of those from W1 as expected.

For LVL, we have compared our estimates with those from Cook et al. (2014) who used a constant mass-to-light ratio and the IRAC 3.6 μm luminosity. The Wen-derived stellar masses are on average ~0.25 dex smaller than the Cook et al. (2014) values, with a scatter of 0.09 dex. Skibba et al. (2011) derived stellar masses for the KINGFISH sample according to the formulation of Zibetti et al. (2009) based on optical and H-band colors. We have compared ours derived using Wen et al. (2013) to theirs and find the values from Skibba et al. (2011) are smaller by ~0.5 dex on average, with a 0.3 dex scatter.

Given the significant uncertainties inherent in the procedures to derive stellar masses over a wide range of galaxy types, and considering the offsets and scales of our new M, estimates, the uncertainties on the stellar masses for the local samples are at most a factor of 2 (0.3 dex). As above, these values are based on a Chabrier (2003) IMF.

2 The distinction used by Wen et al. (2013) is based on colors which are not available for all our samples.
We identified 14 samples at individual galaxies were not available in the required redshift range. We increased statistics when tabulations of observations for in analyses have been avoided where possible, and are used only for the MEGA dataset are also shown as vertical dotted lines. The colors of the portions of the histogram are arbitrary, with the aim of illustrating “low” redshift (blue), “intermediate” redshifts (green), and “high” redshifts (red).

2.2 SDSS10 $z \simeq 0$ galaxy sample

Mannucci et al. (2010) analyzed a set of emission-line galaxies from the SDSS, using the stellar masses from Kauffmann et al. (2003), and SFRs measured from Hα after correcting for extinction by using the Balmer decrement; they reported all values to a Chabrier (2003) IMF. The parameter range covered by this sample is much more limited than the MEGA sample: $9.2 \lesssim \text{dex}(M_*) \lesssim 11.3 M_\odot$; $8.5 \lesssim \log(O/H) \lesssim 9.1$ (assuming the Kewley & Dopita 2002 calibration, see below); $-1.3 \lesssim \log(SFR) \lesssim 0.8 M_\odot \text{yr}^{-1}$. Nevertheless, we include this sample, hereafter SDSS10, in our analysis because of its superb statistics for comparison both locally and at $z > 0$.

2.3 $z > 0$ samples

Because our analysis is focused on observationally constraining metal content at high redshift, to construct the MEGA dataset we have culled from the literature all available samples at $z > 0$ with measured $M_*$, SFR, and O/H. Stacked analyses have been avoided where possible, and are used only to increase statistics when tabulations of observations for individual galaxies were not available in the required redshift range. We identified 14 samples at $z > 0.1$ (see Table 1) for which these three parameters were measured. Unavoidably, this compilation is subject to a variety of selection effects which change with sample and redshift. Nevertheless, from the observational point of view, the MEGA dataset constitutes a unique tool with which to assess basic trends among $M_*$, SFR, and O/H, and establish how they vary with redshift. Table 1 lists the samples that comprise the MEGA dataset, together with their redshift range, selection technique, and other information. We postpone the important discussion of metallicity estimates to Sect. 3.

2.3.1 $0.1 \leq z \leq 0.9$

The most important representative samples in the redshift range $0.1 \leq z \leq 0.9$ come from two surveys, zCOSMOS (Lilly et al. 2009; Cresci et al. 2012) and NewHa (de los Reyes et al. 2015); these two datasets alone comprise 477 galaxies. The first, from COSMOS, is I-band selected and was first described by Lilly et al. (2009). Stellar masses were derived from fitting spectral energy distributions (SEDs) of 12 photometric bands, including Spitzer/IRAC data at 3.6–5.8 μm. SFRs were calculated from Hα and Hβ luminosities, after correcting for extinction either via the Balmer decrement (for galaxies with $z \leq 0.49$) or using the extinction estimated from the SED fitting with an appropriate multiplicative factor. The second large sample in this redshift range comes from the NewHa survey, selected from narrow-Hα band images designed to identify emission-line galaxies around $z \approx 0.8$. de los Reyes et al. (2015) calculate stellar masses through SED fitting of eight photometric bands (up to observed frame J band), and estimate SFRs from the Hα images after correcting for the contribution from [Ni] and for extinction.

Unlike Hunt et al. (2012), we do not include in the MEGA dataset the Luminous Compact Galaxies (LCGs) by Izotov et al. (2011) and the “Green Peas” (Amorín et al. 2010); the latter galaxies are selected by bright [Oiii] λ5007 emission in the SDSS r band (Cardamone et al. 2009). LCGs, instead, are defined by requiring an [Oiii] λ4363 detection, large Hβ equivalent width (EW), and a flux limit in Hα. Thus the LCGs are young (because of the high Hβ EW), highly star forming (because of the Hβ flux limit) and metal poor (because of the [Oiii] λ4363 detection). Although highly interesting objects, the Izotov et al. (2011) selection criteria favor young, metal-poor galaxies, and thus are not representative of abundances of typical galaxy populations at those redshifts.

There are ~60 galaxies in the remaining three samples in this redshift range: galaxies selected from a multi-slit narrowband spectroscopic survey with [Oiii] λλ4959,5007, [Oiii] λλ3727,3729 at $z \sim 0.6 - 0.7$ by Henry et al. (2013); [Oiii] λ4363 DEEP2 selected objects at $z \sim 0.7 - 0.9$ by Ly et al. (2015); and galaxies selected from HST-grism observations ([Oiii], [Oii]) by Xia et al. (2012). These three samples are very interesting because of their selection methods which tend to favor less massive galaxies than typical broadband photometry selections. Stellar masses were derived from SED fitting of COSMOS imaging data including IRAC bands (Henry et al. 2013); of 8-band photometry up to $z'$ for the DEEP2 survey (Ly et al. 2015); and of 10-band HST ACS/WFC3 photometry up to F160W (Xia et al. 2012). With the exception of Henry et al. (2013) who used SED fitting to calculate SFRs, Hα and Hβ corrected luminosities were used to infer SFRs.

2.3.2 $z > 0.9$

Most of the galaxies in this redshift range are color-selected Lyman-Break Galaxies (e.g., Steidel et al. 1999). However, the Quenyer et al. (2009) galaxies are selected from the magnitude-limited Mass Assembly Survey with SINFONI in VVDS (MASSIV, Epinat et al. 2009), and the two $z \sim 1$ samples by Shapley et al. (2005a) and Liu et al. (2008) are
Table 1. Characteristics of the individual samples in the MEGA dataset

| Parent sample | Redshift range | Number | Selection criterion | Original O/H calibration | SF method | Reference |
|---------------|----------------|--------|---------------------|--------------------------|-----------|-----------|
|               |                |        |                     | Local Universe           |           |           |
| KINGFISH      | −0.001 − 0.008 | 55     | Representative      | KK04                     | FUV+TIR±  | Kennicutt et al. (2011) |
| LVL           | −0.001 − 0.003 | 138    | Volume-limited      | M91, KK04, Directb       | FUV+TIR±  | Kennicutt et al. (2009) |
| Starburst     | 0.0 − 0.058    | 41     | Representative      | Directc                  | Hα+24μm4d | Engelbracht et al. (2008) |
| BCD           | 0.009 − 0.044  | 23     | Primordial helium   | Directc                  | Hα+24μm4d | Hunt et al. (2010)       |

**Range:** 0.1 ≤ z ≤ 0.9

|          |                |        |                     |                        |           |           |
|          |                |        |                     |                        |           |           |
| COSMOS   | 0.17 − 0.91    | 334    | [OIII] λ4959,5007   | KK04                     | Hα, Hβ    | Cresci et al. (2012) |
|          | 0.62 − 0.69    | 26     | [OIII] λ4363        | Direct                   | Hβ        | Henry et al. (2013)   |
| DEEP2    | 0.71 − 0.91    | 27     | [OIII] λ4363        | Direct                   | Hβ        | Ly et al. (2015)      |
| NewHo    | 0.79 − 0.82    | 143a   | Narrow-band Hα      | T04                      | Hα        | de los Reyes et al. (2015) |
| HST-grism| 0.60 − 2.32    | 11     | [OIII] λ4959,5007, [OIII] λ3727 | KK04 | Hβ      | Xia et al. (2012) |

z > 0.9

|          |                |        |                     |                        |           |           |
|          |                |        |                     |                        |           |           |
| DEEP2    | 1.02 − 1.40    | 9      | R band              | PP04N2                  | Hα        | Shapley et al. (2005a) |
| DEEP2    | 1.02 − 1.40    | 7      | R band              | PP04N2                  | Hα        | Liu et al. (2008)      |
| VVDS     | 1.27 − 1.53    | 6      | [OIII] λ3727       | T04                     | Hα        | Querey et al. (2009)   |
| BX       | 2.11 − 2.43    | 7      | [OIII] λ3727       | PP04N2                  | Hα        | Shapley et al. (2004)  |
| KBSS     | 2.02 − 2.55    | 79     | Lyman-break dropout | PP04N2                  | Hα        | Steidel et al. (2014)  |
| LSD      | 2.93 − 3.41    | 8      | Lyman-break dropout | KD02                    | Hα        | Mannucci et al. (2009) |
| AMAZE    | 3.04 − 4.87    | 26     | Lyman-break dropout | KD02                    | Hα        | Troncoso et al. (2014) |
| COSMOS   | 2.97 − 3.69    | 35     | Predicted Hβ       | KD02                    | UV        | Onodera et al. (2016)  |

|          |                |        |                     |                        |           |           |
|          |                |        |                     |                        |           |           |
| SXDS/UDS | 1.27 − 1.52    | 5b     | K band              | PP04N2                  | Ho        | Yabe et al. (2014)     |
| COSMOS   | 1.40 − 1.70    | 10b    | sBzK                | PP04N2                  | Ho        | Zahid et al. (2014)    |

|          |                |        |                     |                        |           |           |
|          |                |        |                     |                        |           |           |

**Stacked samples**

|          |                |        |                     |                        |           |           |
|          |                |        |                     |                        |           |           |

|          |                |        |                     |                        |           |           |
|          |                |        |                     |                        |           |           |

— **Notes:**

- If not available, then SFR(FUV), or as last choice SFR(TIR).
- Taken from Berg et al. (2012) or Marble et al. (2010) when available, otherwise from Moustakas et al. (2010) (KK04).
- Taken from Berg et al. (2012), Guseva et al. (2003a), Guseva et al. (2003b), Guseva et al. (2011), Guseva et al. (2012), Izotov & Thuan (2004), Izotov et al. (2006), Izotov et al. (2007), Izotov et al. (2009), Izotov et al. (2012), Kobulnicky & Skillman (1996), Kobulnicky & Skillman (1997), Kniazev et al. (2003), Kniazev et al. (2004), Mattsson et al. (2011), Pérez-Montero & Díaz (2005), Roenback & Bergvall (1995), Shi et al. (2005), Thuan & Izotov (2005), Vigroux et al. (1987), Zhao et al. (2010).
- If not available, then the maximum of SFR(TIR) and SFR(Hα).
- AGN have been excluded.
- This is only one of several “layered” criteria for selecting the galaxies for KBSS Keck-MOSFIRE observations.
- SFRs are taken from Steidel et al. (2014a).
- These are from stacked spectra, but are treated here as individual measurements; the redshifts are taken as the average given in the respective papers (z ~ 1.4, and z ~ 1.6, for Yabe et al. 2014; Zahid et al. 2014, respectively).

selected from the DEEP2 Galaxy Redshift Survey (Davis et al. 2003). Wavelength coverage for stellar-mass determinations varies, with *UBVRIZJK* (Shapley et al. 2004); *BRIKR* (Shapley et al. 2005a; Liu et al. 2008); *UBVRIZJK* (Queyrel et al. 2009); and 14 spectral bands from GOODS-MUSIC (Grazian et al. 2006), including IRAC 3.6, 4.5 μm (Mannucci et al. 2009; Troncoso et al. 2014). Stellar masses for the Onodera et al. (2016) COSMOS sample are fit with *UBVRIZJK* and IRAC bands. Onodera et al. (2016) prefer SFRs inferred from extinction-corrected UV luminosities, but all other SFRs in this redshift range are determined from Hα suitably corrected for extinction.

To ensure better coverage of the redshift range 1.3 < z < 1.7, we have included also the two samples by Yabe et al. (2014) and Zahid et al. (2014). Neither group publishes data for individual galaxies, so we have adopted the parameters of their stacked spectra here as individual galaxies, and used the average redshifts of z ~ 1.4 and z ~ 1.6 for Yabe et al. (2014); Zahid et al. (2014), respectively.

The redshift distribution of the MEGA dataset is shown in Fig. 1, together with the 7 redshift bins that will be used throughout the paper. As mentioned above, Table 1 gives the characteristics of the 19 individual samples comprising the MEGA dataset; there is a total of 990 galaxies from z ~ 0 to z ~ 3.7 (and LnA1689 ~ 2 in the AMAZE sample at z = 4.87).

### 3 METALLICITY CALIBRATIONS

Oxygen abundance O/H is typically used as a proxy for metallicity in emission-line galaxies. Because the ionized gas...
in H\textsc{ii} regions at lower metal abundance is hotter (as measured by electron temperature, $T_e$), the preferred technique to establish O/H is to measure $T_e$ and the physical conditions in the ionized plasma. In this “direct-temperature” or “$T_e$” method, the $T_e$ of the ionized gas is derived from the ratio of the [O\textsc{iii}] $\lambda 4363$ auroral line to lower-excitation lines ([O\textsc{iii}] $\lambda 4959, 5007$); such flux ratios are sensitive to temperature because the auroral and strong lines originate from different excitation states (second and first excited states, respectively). Because the oxygen transitions are collisionally excited, the relative population of the excited states depends on $T_e$. Thus, the strengths of these forbidden lines, combined with the measurement of $T_e$ and density in the nebula, can be converted to an abundance, relative to hydrogen, after correcting for unseen phases of ionization (e.g., Osterbrock & Ferland 2006).

Although the $T_e$ method is more directly related to metallicity, the auroral lines are weak and often difficult to detect, especially at high metallicity. Thus, “strong-line” methods are more generally used to estimate O/H, especially in metal-rich objects and at high redshift. It is necessary to calibrate these methods, either using theoretical photoionization models (e.g., Kewley & Dopita 2002, hereafter KD02), or measurements of $T_e$ (e.g., Pettini & Pagel 2004, hereafter PP04), or a combination of the two (e.g., Denicoló et al. 2002, hereafter D02). Despite the best efforts to correctly cross calibrate these methods over a wide range of physical conditions, there remain large discrepancies, as high as 0.6 dex in log(O/H) (e.g., Kewley & Ellison 2008, and references therein). Thus to correctly assess metal content and its evolution with redshift, it is necessary to apply a common metallicity calibration to the samples under discussion.

In nearby galaxies where spectra can be obtained with sufficient signal-to-noise, the $T_e$ method is generally used. As mentioned above, the most widely used auroral $T_e$ diagnostic line is [O\textsc{iii}] $\lambda 4363$ because of its relative ease of observation, high abundance of emitting ions, and notable strength in the low- and intermediate-metallicity regime (i.e., below solar metallicity). However, there are several potential problems with the $T_e$ method based on [O\textsc{iii}]:

i) Metallicities derived from collisionally-excited lines (CELS) such as [O\textsc{iii}] can be underestimated when temperature fluctuations inside the nebula are present but neglected. The assumption of a single average CEL temperature for the whole nebula, usually higher than the temperature derived from the Balmer discontinuity, tends to lead to an underestimate of the abundances (e.g., Peimbert 1967; Stasińska 2005; Bresolin 2007; Pérez-Montero et al. 2010; Peña-Guerrero et al. 2012).

ii) Additional problems also plague the $T_e$ method including possible non-Boltzmann electron distributions (e.g., Nicholls et al. 2012; Binette et al. 2012; Nicholls et al. 2013); depletion of oxygen onto dust grains (e.g., Peimbert & Peimbert 2010; Peña-Guerrero et al. 2012); and potential shock waves within the nebulae (e.g., Binette et al. 2012).

iii) Finally, recent results suggest that metallicities derived from [O\textsc{iii}] may be more unreliable than those from other auroral lines such as [S\textsc{ii}] $\lambda 6312$ and [N\textsc{ii}] $\lambda 5755$ (e.g., Berg et al. 2015); however, these lines are even more difficult than [O\textsc{iii}] to measure in distant galaxies.

An alternative to the use of $T_e$-diagnostic lines can be the derivation of abundances from optical recombination lines (ORLs), because of their reduced emissivity dependence on density and temperature. Abundances derived from the ratio of the intensity of ORLs tend to be systematically higher than those from CELs (e.g., Peimbert et al. 1993; Liu et al. 1995, 2001; Tsamis et al. 2004; García-Rojas & Esteban 2007). However, such differences may arise from the relation of the ORL abundances to small H-deficient portions of the regions, while the CEL-based metallicities are more representative of the whole nebula (see, e.g., Liu et al. 2000).

Moreover, such lines are extremely faint, thus requiring very high signal-to-noise spectra that are currently available only for the Galaxy and the Local Group (e.g., Blanc et al. 2015).

There are also, perhaps more severe, problems with “strong-line” methods, and the simplifying assumptions made for photoionization model calibrations (e.g., photoionization structure, geometry, stellar age: see Moustakas et al. 2010, for a thorough discussion). As for the $T_e$ method, there may also be systematic discrepancies due to the metallicity-dependent correction for the depletion of oxygen onto dust grains (e.g., Peimbert & Peimbert 2010). Ultimately, the $T_e$ method (with [O\textsc{iii}]) is generally considered to be the most viable, given the limitations with other techniques.

Thus, to ensure the best possible comparison among different samples that rely on different O/H calibrations, it is advantageous to use the strong-line calibration method that most closely resembles values inferred from oxygen-based $T_e$-method estimations. According to the results of Andrews & Martini (2013), who used a stacking technique to measure the oxygen abundances of $\sim 200 000$ star-forming galaxies from the SDSS to enhance the signal-to-noise ratio of the weak [O\textsc{iii}] $\lambda 4363$ line, there are three such methods: PP04 (both [N\textsc{ii}] and [O\textsc{iii}]+[N\textsc{ii}]-based; hereafter PP04N2, PP04O3N2) and D02. Over the metallicity and M* range covered by their calculations of various strong-line methods (see Fig. 10 of Andrews & Martini 2013), the discrepancies between the $T_e$ method and these three methods are $\lesssim 0.1$ dex in $12 + \text{log(O/H)}$.

Thus, in what follows, wherever there are no direct-$T_e$ estimates, we have applied the transformations given by Kewley & Ellison (2008) to convert the original strong-line O/H calibrations for the MEGA dataset (and SDSS10 sample) to the calibrations by D02 and PP04 (PP04N2, PP04O3N2). As reported in Table 1, the original O/H calibrations include: KD02 (Kewley & Dopita 2002; Cresci et al. 2012; Mannucci et al. 2010; Troncoso et al. 2014); KK04 (Kobulnicky & Kewley 2004; Kennicutt et al. 2011; Henry et al. 2013; Xia et al. 2012); M91 (McGaugh 1991; Marble et al. 2010); PP04N2, PP04O3N2 (Pettini & Pagel 2004; Shapley et al. 2004, 2005a; Liu et al. 2008; Yabe et al. 2012; Zahid et al. 2014; Steidel et al. 2014); and T04 (Tremonti et al. 2004; de los Reyes et al. 2015).
Figure 2. Mass-metallicity relation over redshifts from $z \sim 0$ to $z \gtrsim 3.3$, binned as described in the text. The solid (black) curve (labeled $z \sim 0$) corresponds to the fit to SDSS10 (with stacked $T_e$ metallicity determinations) by Andrews & Martini (2013), and the dotted (grey) curves to the polynomial fits by Maiolino et al. (2008) with the KD02 calibration at $z \approx 0.07, 0.7, z \approx 2.2,$ and $z \approx 3.5$. The O/H calibration for all galaxies is PP04N2 as described in the text. Samples are labeled according to the legend in the lower rightmost panel, except for $z \approx 0$ which are: LVL as small open circles (colors correspond to Hubble types with late types $(T \geq 8)$ as cyan, $5 \leq T < 8$ as blue, $3 \leq T < 5$ as magenta, $T < 3$ as red); KINGFISH as (orange) +; Engelbracht et al. (2008) as (purple) filled triangles; Hunt et al. (2010) as (blue) filled squares. The stacked samples (Yabe et al. 2014; Zahid et al. 2014) at $z \approx 1$ are not plotted.

4 SCALING RELATIONS AND THE FUNDAMENTAL PLANE

The MEGA dataset comprises three parameters (pseudo-observables, as they are not directly observed): nebular oxygen abundance ($12+\log(O/H)$), stellar mass ($M_*$), and SFR. As discussed in the Introduction, these three parameters are mutually correlated, although O/H trends flatten at high $M_*$ and high O/H. Here we discuss the scaling relations of the three parameters: the mass-metallicity relation, MZR, the “main sequence” of star formation, SFMS, and the correlation (at least at $z \sim 0$) between sSFR and metallicity.

The MZR with the PP04N2 O/H calibration for different redshift bins is shown in Fig. 2. The solid curve shows the $T_e$-method MZR derived by Andrews & Martini (2013) which well approximates the MEGA dataset at $z \approx 0$. The dotted grey curves represent the polynomial fits given by Maiolino et al. (2008) for the KD02 calibration; at high $M_*$, these curves fail to capture the $T_e$-derived (or PP04N2) metallicities because of the different O/H calibration. As virtually all previous work suggests, the different panels illustrate that as $z$ increases, at a given $M_*$, metallicity decreases. However, at $z \approx 0$ for a given $M_*$, the starburst and BCD samples tend to be more metal-poor than the LVL and KINGFISH galaxies; they behave more like galaxies at $z \gtrsim 1$ than like galaxies in the Local Universe, presumably because of their higher sSFR.

The high sSFRs in the starburst and BCD $z \approx 0$ samples are more clearly seen in Figure 3, which shows the SFMS, or sSFR plotted against $M_*$. The solid line shows the SFMS calibrated with the LVL+KINGFISH samples, having a slope of $-0.19 \pm 0.02$, roughly consistent with that $(-0.23)$ found by Elbaz et al. (2007) for $z \sim 0$ galaxies. The dashed grey lines correspond to the Speagle et al. (2014) formulation for SFR as a function of cosmic time (we have calculated cosmic age for representative redshifts and plotted the result). The slope by Speagle et al. (2014) at $z \gtrsim 2$ is similar to what we find for the Local Universe, which however is shallower (steeper in SFR-$M_*$ space) than their value for $z \sim 0$. Fig. 3 illustrates that as redshift increases, for a given $M_*$, sSFR also increases; galaxies that would be main-sequence galaxies at $z \gtrsim 1$ are starbursts if found at $z \approx 0$. However, it is also seen from the figure that the individual high-$z$ samples do not clearly follow the SFMS; this is almost certainly due to selection effects and will be further discussed in Sect. 5.2. Because of the difficulty in measuring metallicities in high-$z$ emission-line galaxies, flux limits for spectroscopy impose a commensurate limit in SFRs.

The third correlation between sSFR and O/H is shown in Figure 4. As in previous figures, the solid line gives the local calibration on the LVL+KINGFISH and the dotted grey lines show the redshift trend for O/H expected for the higher SFR as predicted by the FPZ (see Sect. 4.2). The SFRs of the local starbursts and BCDs are higher at a given O/H, relative to the other local samples; again, they are more similar to galaxies at $z \gtrsim 1$ than to typical local populations. Sim-
Figure 3. Specific SFR vs. $M_*$ (main sequence of star formation) over redshifts from $z \sim 0$ to $z \gtrsim 3.3$, binned as described in the text. The solid (black) curve (labeled $z \sim 0$) corresponds to the fit to the LVL+KINGFISH samples, and the dotted (grey) curves to the formulation of dependence with $M_*$ and $z$ by Speagle et al. (2014) for $z \approx 0.6$, $z \approx 1.3$, $z \approx 2.3$, and $z \approx 3.3$. Symbols are as in Fig. 2, and the stacked samples (Yabe et al. 2014; Zahid et al. 2014) are not plotted.

Figure 4. Specific SFR vs. $12 + \log(O/H)$ over redshifts from $z \sim 0$ to $z \gtrsim 3.3$, binned as described in the text. The solid (black) line corresponds to the fit to the best-fit slope for the MEGA sample at $z \sim 0$, and the dotted (grey) lines to the prediction of the redshift variation assuming the increase in sSFR given by Speagle et al. (2014) (from $z \sim 0.6$, $z \sim 1.3$, $z \sim 2.3$, $z \sim 3.3$) of the Fundamental Plane for these two parameters described in Sect. 4.2. As in Fig. 2, the O/H calibration for all galaxies is PP04N2 as described in the text. Symbols are as in Fig. 2, and the stacked samples (Yabe et al. 2014; Zahid et al. 2014) are not plotted.
Figure 5. SDSS10 galaxies: $12+\log(O/H)$ plotted against (log of) $M_\ast$. The three panels correspond to number densities of $12+\log(O/H)$ (with the three O/H calibrations) and $M_\ast$. In each panel, the solid curves correspond to the mass-metallicity relation taken from Andrews & Martini (2013) for stacked O/H direct-temperature metallicity determinations as also shown in Fig. 2. The dashed curves show instead the relation of the form used by Andrews & Martini (2013) but fit to the SDSS10 data used here.

Figure 6. SDSS10 galaxies: sSFR plotted against $M_\ast$. The color scale corresponds to the number densities of the two parameters. The solid curve gives the linear SF “main sequence” regression to the LVL+KINGFISH samples (also shown in Fig. 3), and the dashed curve to the Schechter function fitted by Salim et al. (2007).

Similarly to the behavior of the MEGA dataset for the SFMS, the LVL+KINGFISH galaxies show a well-defined correlation between sSFR and O/H, but the correlation disappears for the higher-redshift samples.

The main point of this third correlation is that, at least locally, the three pseudo-observables, $12+\log(O/H)$, SFR, and $M_\ast$ are mutually interdependent. This makes it difficult to determine which is the primary parameter(s) driving the relations, and it is this point which we explore below in Sect. 4.2.

4.1 The SDSS10 relations

Similar correlations are found for the SDSS10 galaxies, although the range in $M_\ast$, O/H, and sSFR is smaller than in the MEGA dataset. Nevertheless, over the limited parameter range the sheer number statistics afford precise determinations of scaling relations and fitting functions which will be important for constraining our models.

Fig. 5 gives the MZR for the three O/H calibrations of SDSS10 sample, transformed from the original KD02. The solid curves, also shown in Fig. 2, give the MZR for the direct-method O/H as found by Andrews & Martini (2013), while the dashed ones are functions of the same form but fit to the SDSS10 dataset itself. The PP04 calibrations (middle and right panels) are, on average, the best approximation to the direct-method O/H curve, although at $12+\log(O/H) \lesssim 8.5$, the D02 calibration is superior. In any case, the functional MZR form used by Andrews & Martini (2013) does not well approximate the SDSS10 data at low mass or low metallicities. The low-mass, low-metallicity linear portion of the data has a slope of $\sim 0.38 \pm 0.003$, similar to the MZR curve, but the latter is offset to higher masses.

The SFMS of the SDSS10 data is shown in Fig. 6; the solid line corresponds to the linear regression for the LVL+KINGFISH galaxies (also shown in Fig. 3) and the dashed curve to the Schechter-like functional form fitted by Salim et al. (2007). Although this last captures the low- and high-mass ends of the SDSS10 data, it does not pass through the region with the highest density (blue colors in Fig. 6). This could have something to do with the different ways that SFR is calculated; Salim et al. (2007) used FUV while Mannucci et al. (2010) used extinction-corrected Hα. Nevertheless, the LVL+KINGFISH regression well approximates this behavior, implying that the SFR derivation is probably not the cause of the discrepancy. The third correlation, between sSFR and $12+\log(O/H)$ is not shown for the SDSS10 data; it shows a similar behavior to the MEGA sample.
4.2 A planar approximation to scaling relations
At high M∗ and O/H, both the MZR and SFMS in- 
flect and flatten (e.g., Tremonti et al. 2004; Noeske et
al. 2007; Whitaker et al. 2014; Lee et al. 2015; Gavazzi
et al. 2015). However, for M∗ below a certain threshold,
M∗ < 3 × 10^10 M⊙, roughly the “turn-over mass” (Tremonti
et al. 2004; Wyder et al. 2007), the relations among the
variables are approximately linear. We propose that, at high M∗
and O/H, the inflections the MZR and the SFMS compens-
sate one another, and hypothesize that even above this in-
fection threshold, the trends in M∗, O/H, and SF can be
approximated by linear relations. Consequently, as discussed
above, these observationally-defined variables could de- 
cine a plane which, given the relatively large scatters in the SFMS,
the MZR, and the SFR-O/H relation, is not well-defined in the
best projection. Because the three parameters are mutually
 correlated, it is important to determine which of the three
is the most fundamental, and whether or not the planar ap-
proximation is sufficient to describe the data. This can be
readily accomplished through a Principal Component Anal-
ysis (PCA, e.g., Hunt et al. 2012).

The MEGA dataset is a significant improvement on the
sample studied by Hunt et al. (2012), and is particularly well
suited for such an analysis. In particular, the MEGA dataset
triples the number of galaxies at z > 2–3 with respect to
Hunt et al. (2012). It spans almost two orders of magnitude in
metallicity (12 + log(O/H) = 7.1 to ~ 9), a factor of ~ 10^6
in SFR (∼ 10^{-3} ≤ SFR ≤ ∼ 10^5 M⊙ yr^{-1}), and a factor of
∼ 10^5 in stellar mass (∼ 10^6 ≤ M∗ ≤ ∼ 10^{11} M⊙); moreover
it includes galaxies at redshifts from z = 0–3.8 (see Fig. 1).
Other samples previously analyzed to find scaling relations
cover much smaller parameter ranges; typically less than a
decade in metallicity (12 + log(O/H) > 8.4), a factor of ~200
in SFR (∼ 0.04 ≤ SFR ≤ 6.0 M⊙ yr^{-1}), and roughly 2 orders of
magnitude in stellar mass (M∗ ≥ 10^8 M⊙) (e.g., Tremonti et
al. 2004; Mannucci et al. 2010; Lara-López et al. 2010; Yates
et al. 2012). Because more than 50% of the MEGA dataset
has z > 0.5, and it includes galaxies at redshift z ≥ 3.5,
we can test the assumption that the relations among the
observationally-defined variables are redshift invariant.

4.3 PCA of the MEGA and SDSS10 samples
We have therefore performed a PCA for all three O/H cal-
bribations of the MEGA dataset without imposing a limit in M∗
(c.f., Hunt et al. 2012). A PCA diagonalizes the 3D
covariance matrix, thus defining the orientation of the pa-
rameter space which minimizes the covariance. The orienta-
tion is contained in the eigenvectors which are, by definition,
mutually orthogonal. If the 3D space formed by the three
pseudo-observables is truly planar, we would expect most
of the variance to be contained in the first two eigenvectors
(the orientation of the plane): for the third eigenvector, per-
pendicular to the plane, the variance should be very small.

Independently of the O/H calibration, the PCA shows
that the set of three observables truly defines a plane: ≥98%
of the total variance is contained in the first two eigene-

vectors. Most (87%) of the variance is contained in the first

eigenvector alone (or Principal Component, PC), PC1; it
is dominated by SFR, with M∗ contributing slightly less,
and O/H giving only a marginal contribution. PC2, the sec-
ond eigenvector, holds 10–11% of the variance, and is dom-
inated by M∗, followed by SFR, and as in PC1, with O/H
again only marginal. The smallest fraction of the variance
(~ 1.5–1.8%) is contained in the third eigenvector, PC3,
which is dominated by O/H; the implication is that O/H is
the most dependent parameter, governed almost completely
by M∗ and SFR. Moreover, this means that the 3D space
defined by O/H, SFR, and M∗ is degenerate; because of
the mutual correlations of the pseudo-observables, only two
parameters are required to describe the properties of the
galaxies.

We have also performed a PCA of the SDSS10 galax-
ies, and obtained similar results: namely, the third eigene-
ctor, PC3, the one dominated by O/H, contains the smallest
fraction of the variance. The residuals of 0.05–0.06 dex are
comparable to that obtained by the FMR formulation by
Mannucci et al. (2010).

That star-forming galaxies form a plane in O/H, SFR,
and M∗ is not a new result. Lara-López et al. (2010) con-
cluded that the 3D space of O/H, SFR, and M∗ of ~ 33,000
SDSS galaxies could be represented as a plane by a prin-
cipal component analysis rather than a PCA (although see
Lara-López et al. 2013). Hunt et al. (2012) derived a PCA for a
sample similar to ours, although dominated by LCGs, and
also concluded that a 2D plane was sufficient to describe the
3D dataset.

As in the Introduction, we will refer to the resulting 2D
plane as the FPZ (Fundamental Plane in metallicity). The
FPZ for the MEGA dataset is shown in the top panel of Fig-
ure 7 where we have plotted 12 + log(O/H) vs. the equa-
tion that results from equating PC3 (PP04N2) to zero (see Table
2 for the other O/H calibrations):

\[12 + \log(O/H) = -0.14 \log(SFR) + 0.37 \log(M_\star) + 4.82 \]  

The bottom panel of Fig. 7 shows the residuals from the FPZ
for the different O/H calibrations. For PP04N2, they are well
approximated by a Gaussian with a σ = 0.16, corresponding
to ≤ 45% uncertainty; the other two O/H calibrations (D02,
PP04O3N2) give similar results, although slightly larger (see Table
2). The residuals of the FPZ relation are independent of
redshift to within 0.16 dex, the overall uncertainty; nev-

ertheless the different symbols plotted in the top panel
(and the different histograms in the bottom one) suggest some
slight deviation with redshift which we will explore in Sect.
5.3.

Because of the turnover of the MZR and SFMS at high
stellar masses, our assumption of linearity in the FPZ could
also produce a residual correlation with M∗. We have in-
vestigated this possibility and found that the FPZ resid-
uals and (log)M∗ are uncorrelated; the mean residuals of
the regression are ~0.16 dex (for the PP04 calibration), the
same as those of the FPZ itself. Moreover, the slope of the
FPZ residuals vs. (log)M∗ is zero to within the uncertainties
(~0.015 ± 0.01). Thus, our hypothesis that the inflections
in the MZR and SFMS compensate one another is appar-
tently justified; the curvature in the MZR can be ade-
quately approximated by M∗ and SFR. Moreover, the slope
of the MZR and SFMS compensation one another is appar-
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(0.1 dex), and also higher than the FMR (0.06 dex) found for the SDSS10 sample by Mannucci et al. (2010). However, as shown in Table 2, the SDSS10 data span limited ranges in O/H, SFR, stellar mass (and redshift) relative to the parameter space covered by the MEGA data. The mean and standard deviation of (log) stellar mass for the SDSS10 sample is $10.26 \pm 0.41 M_\odot$, while the comparable mean, standard deviation for MEGA is $9.44 \pm 0.93 M_\odot$; (log) SFR shows a similar pattern: $-0.04 \pm 0.43 M_\odot$ yr$^{-1}$ for SDSS10 compared to $0.16 \pm 1.12 M_\odot$ yr$^{-1}$ for MEGA. Thus, the higher dispersion in the MEGA FPZ is not surprising despite the many more galaxies in SDSS.

The value of the FPZ dispersion is lower than the scatter of the MZR for $\sim 20,000$ VVDS galaxies within individual redshift bins from $z \sim 0.3 - 0.9$ ($\sim 0.20$ dex, Lamareille et al. 2009). The FPZ dispersion for the $\sim 1000$ galaxies studied here is only slightly higher than that found for the MZR of 25 nearby dwarf galaxies (0.12 dex, Lee et al. 2006), a sample dominated by low-mass galaxies. It is also only slightly higher than the rms scatter of 0.12 dex found by Henry et al. (2013) for 18 galaxies at $z \sim 0.6$ in the mass range dex(8.5)$< M_\star <$dex(9.0). Because the dispersion in the MZR is found to increase with decreasing $M_\star$ (Tremonti et al. 2004; Mannucci et al. 2011), a $\sigma$ of 0.16 dex is a reasonable value, given the broad parameter space covered by our dataset.
4.4 Comparison with the FMR

Table 2 gives the mean residuals and offsets of the FPZ applied to the MEGA and SDSS10 datasets, as well as of the FMR from Mannucci et al. (2010) applied to SDSS10, and the FMR extended to lower M∗ by Mannucci et al. (2011) applied to the MEGA dataset. Despite the vastly different parameter ranges over which the FPZ and FMR are calibrated, results from the Table show that the FPZ and the FMR are roughly equivalent in terms of the width of the residuals, i.e., the accuracy of the approximation. The FPZ fits the SDSS10 dataset almost as well as the FMR itself (σ ≈ 0.08 dex), and the FMR is reasonably good at reproducing the metallicities of the MEGA dataset.

However, the salient difference between the FPZ and FMR formulations is the negative O/H offsets of the FMR; Col. (4) of Table 2 shows that the FMR predicts metallicities both for the MEGA dataset and for the recalibrated SDSS10 that can be in excess by as much as ~ −0.3 dex. A similar result was found by Hunt et al. (2012) relative to the FMR, although the offset was larger, ~ −0.4 dex, presumably because of the original SDSS10 KD02 calibration.

Indeed, the over-large metallicities predicted by the FMR are almost certainly due to the different O/H calibrations as discussed in Sect. 3. Cullen et al. (2014) found a similar discrepancy of their observations at z ~ 2 with respect to the FMR using the same (KD02) calibration as Mannucci et al. (2010). Some groups concluded that the FMR evolves with redshift because of its failure to fit galaxies at z ~ 2 − 3 (e.g., Steidel et al. 2014; Troncoso et al. 2014). However, we find that the FPZ, unlike the FMR, is apparently invariant with redshift; our result almost certainly stems from the common O/H calibration and its similarity to the Te method by which strong-line methods are calibrated at low metallicity. Thus, it is of extreme importance to compare galaxies at different redshifts with a common O/H calibration that is as accurate as possible at low metallicities, and that smoothly connects these with the difficult intermediate-metallicity regime and with higher metallicities nearer to or exceeding Solar.

5 METALLICITY AND SFR COEVOLUTION

Much work has been done to establish how metal abundance and SFR vary with redshift. As mentioned in the Introduction, the picture that emerges from these studies is that the shape of the MZR is relatively invariant while the metallicity for a given M∗ decreases with increasing redshift (e.g., Shapley et al. 2005a; Cowie & Barger 2008; Henry et al. 2013; Yabe et al. 2014; Erb et al. 2006a; Zahid et al. 2012; Steidel et al. 2014; Maiolino et al. 2008; Mannucci et al. 2009; Troncoso et al. 2014; Onodera et al. 2016). At the same time, it is well known that the SFMS also remains relatively constant in shape, but at a given M∗, SFMR (and sSFR) increases with redshift (e.g., Noeske et al. 2007; Karim et al. 2011; Speagle et al. 2014). In the context of the FPZ, the relatively small dispersion of the residuals suggests that the FPZ formulation is apparently invariant with redshift to z ~ 3.7, even with the new MEGA sample that more than triples the number of galaxies at z ≥ 2 − 3 with respect to Hunt et al. (2012). Thus, under the hypothesis that the FPZ is maintained even at high z, the opposing redshift trends of O/H and SFR must somehow be mutually compensated. For typical galaxy populations, at fixed M∗, the increase of SFR with redshift must be accompanied by a corresponding decrease in O/H. We can quantify such trends with the MEGA dataset and the FPZ. Table 3 reports the median values of M∗, 12+log(O/H), and sSFR for the MEGA dataset for 6 mass bins within the 7 redshift bins shown in Fig. 1.

5.1 Redshift variation of O/H and sSFR

The trends with redshift of the MEGA dataset are shown in Fig. 8 where 12+log(O/H) and sSFR are plotted vs. redshift z. Following Karim et al. (2011), we have fit the redshift variation with separable functions in M∗ and z:

\[ \text{sSFR, O/H}(M_\ast, z) \propto M_\ast^{\beta} (1 + z)^{n} \]

The dependence on stellar mass is encompassed in the power-law index β, and the z dependence in the power-law index (slope in log space). n. The symbols in Fig. 8 correspond to data binned in redshift and in stellar mass as given in the legend (see Fig. 1 for redshift intervals); the error bars give the 25% quantiles of the data within each bin.

For the trend of sSFR \( \propto (1 + z)^n(sSFR) \), the mean power-law index \( \langle n(sSFR) \rangle = 2.8 \pm 0.5 \). For the M∗ bins between dex(8.5) and dex(10) M⊙, n is relatively constant: 3.1 ± 0.37. This value is roughly consistent with n(sSFR) ~ 3.4 − 3.5 to z ~ 2 as reported by Oliver et al. (2010); Karim et al. (2011). For larger M∗ (M∗ ≥ 10^10 M⊙), the index decreases to \( n(sSFR) = 2.2 \pm 2.4 \). Similar slopes and such a flattening are also seen in the highly star-forming sub-sample of COSMOS galaxies described by Karim et al. (2011); the MEGA dataset probably represents a similarly highly star-forming sample, at least at the higher redshifts.

The redshift variation of O/H is represented by the mean power-law index \( \langle n(O/H) \rangle \), averaged over all mass bins (except the lowest one, because of the lack of low-mass galaxies at z ≥ 1) is: \( \langle n(O/H) \rangle = 2.8 \pm 0.5 \). For the M∗ bins between dex(8.5) and dex(10) M⊙, n is relatively constant: 3.1 ± 0.37. This value is roughly consistent with n(O/H) ~ 3.4 − 3.5 to z ~ 2 as reported by Oliver et al. (2010); Karim et al. (2011). For larger M∗ (M∗ ≥ 10^10 M⊙), the index decreases to \( n(O/H) = 2.2 \pm 2.4 \). For 12+log(O/H) ~ (1 + z)^n(O/H) (left panel of Fig. 8). From the relatively small standard deviation, and visually evident in Fig. 8, it is apparent that the index n(O/H) is much more constant over variations in M∗ than the equivalent index n(sSFR) for sSFR. Table 4 gives the fitted coefficients for the (PP04N2) metallicity redshift variation of the MEGA dataset; these are the equations describing the dashed curves in the left panel of Fig. 8 and subsequent figures.

5.2 Variation of O/H and sSFR with stellar mass

We now examine the M∗ dependence of the redshift variations of O/H and sSFR. Fig. 9 shows 12+log(O/H) vs. M∗ (left panel) and sSFR vs. M∗ (right), binned into different redshift bins; these are equivalent to the changes with redshift of the MZR (left panel) and the SFMS (right).

Here we assess β in the separated formalism as above: sSFR, O/H \( \propto M_\ast^{\beta} (1 + z)^{n} \). β(O/H) corresponds roughly to the slope (power-law index) of the MZR, and β(sSFR) to the slope of the SFMS. For the M∗ trend of O/H, \( \langle \beta \rangle \) (averaged over all mass bins) is: \( \langle \beta \rangle (O/H) = 0.21 \pm 0.05 \). Both the normalization and the slope of the MZR
Table 2. FPZ and FMR applied to MEGA and SDSS datasets

| Sample    | Calibration | Offset(fit)b | (12+log(O/H))²c | (Log(SFR))²c | (Log(M₉))²c | 12+log(O/H) =     |
|-----------|-------------|--------------|------------------|--------------|-------------|-----------------|
|           | (1)         | (2)          | (3)              | (4)          | (5)          | (6)             | (7)             | (8)            |
| MEGA D02  | 0.185       | 0.03         | 8.422 ± 0.33     | 0.159 ± 1.12 | 9.440 ± 0.93 | −0.17 s + 0.44 m + 4.33 |
| MEGA PP04N2 | 0.159      | 0.02         | 8.332 ± 0.30     | 0.150 ± 1.12 | 9.440 ± 0.93 | −0.14 s + 0.37 m + 4.82 |
| MEGA PP04O3N2 | 0.178     | 0.03         | 8.355 ± 0.32     | 0.159 ± 1.12 | 9.440 ± 0.93 | −0.16 s + 0.41 m + 4.48 |

SDSS10 D02 | 0.102       | −0.04        | 8.768 ± 0.09     | −0.038 ± 0.43 | 10.263 ± 0.41 | As above for D02. |
| SDSS10 PP04N2 | 0.080     | −0.004       | 8.660 ± 0.11     | −0.038 ± 0.43 | 10.263 ± 0.41 | As above for PP04N2. |
| SDSS10 PP04O3N2 | 0.088     | −0.01        | 8.709 ± 0.12     | −0.038 ± 0.43 | 10.263 ± 0.41 | As above for PP04O3N2. |

The slope β(2014) can be attributed to their use of a Salpeter IMF, which is rather constant up to β ∼ 2; β = 0.24 ± 0.03, 12+log(O/H)(M₉ = dex(9) M☉) = 8.28 ± 0.04 (averages and standard deviations over 5 equally-weighted redshift bins for z < 1.8). However, for z > 2, both the normalization and the slope gradually decrease: β = 0.15 ± 0.03 and 12+log(O/H)(M₉ = dex(9) M☉) = 8.01 ± 0.03 within the highest (z ~ 3) redshift bin.

The slope β(O/H) for the MEGA dataset at z ~ 1.4 of 0.23 ± 0.02 is slightly steeper than that by Yabe et al. (2014) who find β = 0.15, but shallower than the slope of 0.3 found by Liu et al. (2008) using PP04N2 in a similar redshift range (both samples are also included in the MEGA dataset). The steeper slopes we find relative to Yabe et al. (2014) can be attributed to their use of a Salpeter IMF, rather than the Chabrier (2003) IMF used here. In the redshift range z ~ 0.5–0.9, Cowie & Barger (2008) find an MZR slope of 0.13 – 0.17 using the KK04 and T04 O/H calibrations and a Salpeter IMF; this is also somewhat shallower than the PP04N2 β = 0.24 ± 0.02 in a similar redshift range, but with the Chabrier IMF. The steeper slopes we find for the MEGA dataset are consistent with those found by Zahid et al. (2011) for DEEP2 galaxies at z ~ 0.8 using the KK04 O/H calibration (and a Chabrier IMF). As discussed by Zahid et al. (2011), differences in fitting procedures are an important consideration in comparing slopes of the MZR, but the O/H calibration is also important. Kewley & Ellison (2008) illustrate that both the slope (at the low-mass end) and the absolute O/H determination depend strongly on the calibration. Thus, the consistency of the MZR slopes relative to previous work lends confidence our approach.

For the trends of sSFR with M₉, corresponding to the SFMS, averaging over all redshifts gives β(sSFR) = −0.51 ± 0.24 (average and standard deviation over the individual redshift bins); the mean slope is poorly determined because of the (possibly spurious, see below) steepening toward high z. At z ~ 0, we find β(sSFR) = −0.21 ± 0.025. At z ~ 0.25, we find a steeper slope, β(sSFR) = −0.29 ± 0.07, roughly consistent with the sSFR

3 This slope is derived from all data at z ~ 0, while the slope are rather constant up to z ~ 2; β = 0.24 ± 0.03, 12+log(O/H)(M₉ = dex(9) M☉) = 8.28 ± 0.04 (averages and standard deviations over 5 equally-weighted redshift bins for z < 1.8). However, for z > 2, both the normalization and the slope gradually decrease: β = 0.15 ± 0.03 and 12+log(O/H)(M₉ = dex(9) M☉) = 8.01 ± 0.03 within the highest (z ~ 3) redshift bin.

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3 This slope is derived from all data at z ~ 0, while the slope
| Redshift bin | Mass bin | Number | \( \log(M_*) \) (\( M_\odot \)) | \( 12 + \log(O/H) \) (PP04N2) | \( \log(sSFR) \) (yr\(^{-1}\)) |
|--------------|----------|--------|-------------------------------|-------------------------------|-------------------------------|
| \( z \leq 0.1 \) | \( \log(M_*) < 8.5 \) | 128 | 7.76 ±0.31 | 7.92 ±0.12 | -9.61 ±0.28 |
| | \( 8.5 \leq \log(M_*) < 9 \) | 24 | 8.69 ±0.09 | 8.16 ±0.10 | -9.75 ±0.18 |
| | \( 9 \leq \log(M_*) < 9.5 \) | 34 | 9.18 ±0.11 | 8.39 ±0.13 | -9.84 ±0.18 |
| | \( 9.5 \leq \log(M_*) < 10 \) | 18 | 9.64 ±0.05 | 8.39 ±0.21 | -9.94 ±0.11 |
| | \( 10 \leq \log(M_*) < 10.5 \) | 26 | 10.28 ±0.12 | 8.69 ±0.14 | -9.99 ±0.15 |
| | \( \log(M_*) > 10.5 \) | 27 | 10.71 ±0.11 | 8.75 ±0.03 | -10.36 ±0.62 |
| \( 0.1 < z < 0.4 \) | \( \log(M_*) < 8.5 \) | 1 | 8.09 | 8.16 | -8.91 |
| | \( 8.5 \leq \log(M_*) < 9 \) | 16 | 8.84 ±0.08 | 8.22 ±0.08 | -9.23 ±0.11 |
| | \( 9 \leq \log(M_*) < 9.5 \) | 39 | 9.35 ±0.09 | 8.34 ±0.10 | -9.52 ±0.27 |
| | \( 9.5 \leq \log(M_*) < 10 \) | 84 | 9.71 ±0.11 | 8.53 ±0.11 | -9.60 ±0.30 |
| | \( 10 \leq \log(M_*) < 10.5 \) | 18 | 10.11 ±0.09 | 8.62 ±0.06 | -9.76 ±0.25 |
| | \( \log(M_*) > 10.5 \) | 6 | 10.57 ±0.04 | 8.60 ±0.09 | -9.73 ±0.08 |
| \( 0.4 < z < 0.7 \) | \( \log(M_*) < 8.5 \) | 9 | 8.38 ±0.02 | 8.19 ±0.04 | -8.58 ±0.30 |
| | \( 8.5 \leq \log(M_*) < 9 \) | 12 | 8.82 ±0.13 | 8.25 ±0.05 | -9.02 ±0.36 |
| | \( 9 \leq \log(M_*) < 9.5 \) | 49 | 9.34 ±0.07 | 8.35 ±0.13 | -9.25 ±0.17 |
| | \( 9.5 \leq \log(M_*) < 10 \) | 84 | 9.69 ±0.08 | 8.44 ±0.08 | -9.41 ±0.21 |
| | \( 10 \leq \log(M_*) < 10.5 \) | 12 | 10.12 ±0.10 | 8.59 ±0.11 | -9.37 ±0.12 |
| | \( \log(M_*) > 10.5 \) | 3 | 10.53 ±0.04 | 8.64 ±0.01 | -9.51 ±0.15 |
| \( 0.7 < z < 0.9 \) | \( \log(M_*) < 8.5 \) | 6 | 8.10 ±0.18 | 8.04 ±0.12 | -7.58 ±0.46 |
| | \( 8.5 \leq \log(M_*) < 9 \) | 16 | 8.73 ±0.12 | 8.17 ±0.18 | -8.07 ±0.39 |
| | \( 9 \leq \log(M_*) < 9.5 \) | 45 | 9.34 ±0.11 | 8.33 ±0.12 | -9.43 ±0.43 |
| | \( 9.5 \leq \log(M_*) < 10 \) | 91 | 9.73 ±0.08 | 8.44 ±0.11 | -9.21 ±0.27 |
| | \( 10 \leq \log(M_*) < 10.5 \) | 22 | 10.21 ±0.08 | 8.60 ±0.07 | -9.24 ±0.27 |
| | \( \log(M_*) > 10.5 \) | 20 | 10.74 ±0.04 | 8.69 ±0.04 | -9.54 ±0.34 |
| \( 0.9 < z < 1.8 \) | \( \log(M_*) < 8.5 \) | 0 | - | - | - |
| | \( 8.5 \leq \log(M_*) < 9 \) | 1 | 8.74 | 7.49 | -7.98 |
| | \( 9 \leq \log(M_*) < 9.5 \) | 2 | 9.20 ±0.02 | 8.08 ±0.05 | -8.88 ±0.08 |
| | \( 9.5 \leq \log(M_*) < 10 \) | 14 | 9.81 ±0.17 | 8.41 ±0.06 | -8.81 ±0.17 |
| | \( 10 \leq \log(M_*) < 10.5 \) | 14 | 10.24 ±0.11 | 8.52 ±0.03 | -8.67 ±0.06 |
| | \( \log(M_*) > 10.5 \) | 12 | 10.68 ±0.13 | 8.54 ±0.03 | -9.14 ±0.10 |
| \( 1.8 < z < 2.8 \) | \( \log(M_*) < 8.5 \) | 0 | - | - | - |
| | \( 8.5 \leq \log(M_*) < 9 \) | 2 | 8.74 ±0.07 | 8.12 ±0.01 | -7.24 ±0.07 |
| | \( 9 \leq \log(M_*) < 9.5 \) | 5 | 9.29 ±0.10 | 8.25 ±0.05 | -7.05 ±0.10 |
| | \( 9.5 \leq \log(M_*) < 10 \) | 33 | 9.69 ±0.16 | 8.24 ±0.08 | -7.10 ±0.16 |
| | \( 10 \leq \log(M_*) < 10.5 \) | 24 | 10.18 ±0.12 | 8.28 ±0.09 | -8.76 ±0.26 |
| | \( \log(M_*) > 10.5 \) | 24 | 10.89 ±0.26 | 8.46 ±0.14 | -9.21 ±0.20 |
| \( 2.8 < z \leq 3.8 \) | \( \log(M_*) < 8.5 \) | 2 | 8.41 ±0.02 | 7.93 ±0.01 | -7.90 ±0.26 |
| | \( 8.5 \leq \log(M_*) < 9 \) | 6 | 8.82 ±0.08 | 7.88 ±0.09 | -7.61 ±0.18 |
| | \( 9 \leq \log(M_*) < 9.5 \) | 12 | 9.35 ±0.06 | 8.12 ±0.17 | -8.09 ±0.20 |
| | \( 9.5 \leq \log(M_*) < 10 \) | 24 | 9.71 ±0.14 | 8.16 ±0.15 | -8.19 ±0.22 |
| | \( 10 \leq \log(M_*) < 10.5 \) | 16 | 10.12 ±0.08 | 8.21 ±0.04 | -8.77 ±0.19 |
| | \( \log(M_*) > 10.5 \) | 8 | 10.75 ±0.14 | 8.19 ±0.02 | -8.86 ±0.31 |

\(^a\) Medians of values within each redshift bin; the upper and lower values correspond to the 75\% and 25\% quantile levels, respectively. We have not considered the AMAZE galaxy, LHA1689–2, at \( z = 4.87 \).

vs. \( M_* \), power-law index of \(~ -0.4\) estimated by Karim et al. (2011) and by Speagle et al. (2014) for \( z \approx 0.3 \). At \( z \sim 3 \), \( \beta(sSFR) = -0.64 \pm 0.09 \); the observed steepening of \( \beta(sSFR) \) toward higher redshift evident in Fig. 9 is inconsistent with the results of Speagle et al. (2014) who find of \(-0.19 \pm 0.02\) in Sect. 4 is found from the LVL+KINGFISH galaxies only; the two slopes are in good agreement.

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steeper slopes with increasing redshift in (log) SFR vs. $M_*$, corresponding to shallower slopes in (log) sSFR.

Indeed, the MEGA dataset does not show clear evidence for a SFMS within individual redshift bins (see Fig. 3); this could be because the galaxies at higher redshift are selected basically for a constant SFR (see Table 1) rather than a selection based on $M_*$. Such a selection can result in a basically flat trend of SFR with $M_*$ (e.g., Erb et al. 2006b; Lee et al. 2013; Renzini & Peng 2015), which produces a steep dependence of sSFR with $M_*$, $sSFR \propto M_*^{-1}$, similar to the behavior of the MEGA dataset at high $z$. This is essentially a Malmquist bias since at low stellar masses, only galaxies

Figure 8. Binned measurements of $12+\log(O/H)$ and sSFR as a function of redshift using the MEGA dataset. As in previous figures, the O/H calibration is PP04N2; the mass bins are shown in the lower right-hand corner. Error bars correspond to the 25th percentile of the vertical parameter within each mass bin. The curves are an approximation to the observed trends obtained by adopting the formulation of Karim et al. (2011) based on separable functions of $M_*$ and $z$: sSFR, $O/H \propto M_*^{\beta} (1+z)^n$. The mean power-law indices, $n$, averaged over all mass bins (except the lowest one) are: $\langle n \rangle (O/H) = -0.57 \pm 0.17$ for $12+\log(O/H) \propto (1+z)^n$ (left panel), and $\langle n \rangle (sSFR) = 2.79 \pm 0.52$ for sSFR $\propto (1+z)^n$ (right).

Figure 9. Binned measurements of $12+\log(O/H)$ and sSFR as a function of $M_*$ using the MEGA dataset. As in previous figures, the representative O/H calibration is PP04N2; the redshift bins are shown in the corners of the figures (see also Fig. 1). Error bars correspond to the 25th percentile of the vertical parameter within each redshift bin. As in Fig. 8, the curves are an approximation to the observed trends obtained by adopting the formulation of Karim et al. (2011) based on separable functions with the redshift trend given by $(1+z)^n$, and the $M_*$ variation by $M_*^{\beta}$. The mean power-law indices, $\beta$, averaged over all redshift bins are: $\langle \beta \rangle (O/H) = 0.21 \pm 0.05$ for $12+\log(O/H) \propto M_*^{\beta} (1+z)^n$ (left panel), and $\langle \beta \rangle (sSFR) = -0.51 \pm 0.24$ for sSFR $\propto M_*^{\beta} (sSFR)$ (right).
with relatively high SFR are selected. Because the MEGA dataset requires emission lines in order to measure metallicity spectroscopically, such an effect almost certainly plays an important role (e.g., Juneau et al. 2014). Ultimately, because of such selection effects, the MEGA dataset may not be completely representative of the SFR-M_* correlations at high redshift. Nevertheless, it is the best dataset currently available for assessing the evolution of the MZR.

5.3 Redshift invariance of the FPZ?

One way to assess the redshift invariance of the FPZ is by comparing the coefficients for redshift variation discussed above in Sect. 5.1 with those for the FPZ. We thus performed multi-variable linear regressions on the MEGA dataset for 12+log(O/H) as a function of M_* and redshift, and the same for sSFR. Performing a robust fit\(^4\), we find (for the PP04N2 calibration):

\[
12 + \log(O/H) = 0.27 \log(M_*) - 0.59 \log(1+z) + 5.89 \tag{2}
\]

with a residual standard error of \(\sim 0.15\) dex, and

\[
\log(sSFR) = -0.29 \log(M_*) + 2.88 \log(1+z) - 7.16 \tag{3}
\]

with a residual standard error of \(\sim 0.44\) dex. The FPZ in Eqn. (1) can be expressed as a function of sSFR, rather than SFR:

\[
12 + \log(O/H) = -0.14 \log(sSFR) + 0.23 \log(M_*) + 4.82 \tag{4}
\]

By inserting the redshift variation of sSFR given by Eqn. (3) in Eqn. (4) for the FPZ, and comparing it with the redshift variation of O/H given by Eqn. (2), we can compare the resulting difference equation term by term and assess the redshift invariance of the FPZ; we would expect \(\approx 0\) in such a case. The resulting coefficient for the difference (FPZ - z fits) of the \(\log(M_*)\) term is 0.0004, consistent with 0. For \(\log(1+z)\), we find a difference of 0.178, roughly consistent with 0 to within the residual standard errors of the fits. The resulting difference for the constant term is \(\sim -0.051\), again consistent with 0 within the standard errors. Although this result pertains to PP04N2, similarly small difference coefficients are obtained for the D02 and PP04O3N2 calibrations. \(\textit{We thus conclude that the FPZ is approximately redshift invariant to within 0.15-0.16\,\text{dex}}\) (see Sect. 4.2); for typical galaxy populations the increase of sSFR with redshift is compensated by the decrease in O/H (although see Wuyts et al. 2014).

Nevertheless, the \(\log(1+z)\) difference coefficient of \(\sim 0.18\) dex is slightly larger than would be expected given the residuals of the (PP04N2) FPZ, \(\sim 0.16\) dex. This implies that the FPZ is not a perfect formulation of the redshift evolution of metallicity. Therefore, to investigate the amplitude of the residual trend with redshift, in Fig. 10 we have plotted the FPZ O/H residuals vs. redshift. The trend (shown by the solid line) is significant with:

\[
12 + \log(O/H) - \text{FPZ(PP04N2)} = (-0.048 \pm 0.006) z + 0.047 \pm 0.007 \tag{5}
\]

This implies that at \(z \approx 3.5\), the FPZ predicts metallicities \(12+\log(O/H)\) that are roughly 0.17 dex too large. However, the residual standard error of the (PP04N2) fit in Eqn. (5) is 0.16 dex, equivalent to the residuals of the FPZ itself; indeed, the spread of residuals at \(z \sim 0\) is as large or larger than the spread of residuals at \(z \geq 3\). Thus, while the current data suggest that the FPZ may not fully describe metallicity evolution (or its lack thereof related to SFR), the discrepancies are within the overall noise in the estimation. The data at \(z \geq 3\) are still relatively sparse, however, and more data with accurate metallicity measurements should help in confirming (or refuting) this conclusion.

6 DISCUSSION AND SUMMARY

The FPZ presented in Sect. 4.2 is based on the hypothesis that the curvature in the MZR at high stellar masses is compensated by the inflection in the SFMS. Our results show

\[
\text{Table 4. Fitted (PP04N2) O/H redshift variation of the MEGA dataset}^a
\]

| Mass bin \(\log(M_*)\) | Number of redshift points | \(a\) | \(b\) | Median RMS residual |
|------------------------|--------------------------|-----|-----|------------------|
| 8.5 – 9.0              | 77                       | 0.23 ± 0.03 | -0.33 ± 0.11 | 0.010 ± 0.18 |
| 9.0 – 9.5              | 168                      | 8.43 ± 0.02 | -0.45 ± 0.07 | 0.002 ± 0.15 |
| 9.5 – 10.0             | 349                      | 8.59 ± 0.01 | -0.64 ± 0.05 | 0.002 ± 0.14 |
| 10.5 – 10.5            | 132                      | 8.72 ± 0.02 | -0.75 ± 0.05 | 0.012 ± 0.13 |
| > 10.5                 | 100                      | 8.78 ± 0.02 | -0.68 ± 0.06 | 0.005 ± 0.14 |

\(^a\) Coefficients for robust fits to the individual data points (within each mass bin) of the form \(12+\log(O/H) = a + b \log(1+z)\); \(b\) corresponds to \(n(O/H)\) as described in the text.

\(^4\) For all statistical calculations we use R, a free software environment for statistical computing and graphics, \(\text{https://www.r-project.org/}\).
that this hypothesis is reasonably good, at least to within $\sim 0.16$ dex in $12 + \log(O/H)$. Comparison of the FPZ with the FMR shows that both formulations adequately represent the mutual correlations of $M_*$, SFR, and O/H, but also that the O/H calibration is crucial; applying the FMR to arbitrary samples can result in metallicity offsets as large as $-0.2$ to $-0.3$ dex, compared with the three O/H calibrations considered here.

Some groups have concluded that at a given redshift or over a narrow range of redshifts, that the MZR does not depend on SFR (e.g., Wuyts et al. 2014; Sanders et al. 2015). However, the parameters of the MEGA sample in specific redshift intervals, given in Table 3, are comparable to the other samples used to draw these conclusions. It is likely that the broad parameter space spanned by the MEGA dataset contribute to the differences in the outcome. The PCA analysis presented here could also play a role; indeed, if we fit the MEGA data with a simple multi-variable regression of $12 + \log(O/H)$ with respect to $M_*$, SFR, and redshift, we find very little dependence of O/H on SFR. This is because of the strong dependence of SFR (and sSFR) on redshift through the increasing normalization of the SFMS (e.g., Sect. 5). The mutual correlations of the variables underlying the FPZ must be taken into account for any analysis considering the MZR and its dependence on SFR.

In conclusion, we have compiled a new MEGA dataset consisting of $\sim 1000$ galaxies taken from 19 individual samples spanning a wide range of stellar masses, SFRs, and metallicities and covering redshifts from $z \approx 0$ to $z \approx 3.7$. In addition to larger numbers of high-z galaxies, the main improvement of this dataset over that of Hunt et al. (2012) is the common O/H calibrations derived for the MEGA galaxies. The main results are as follows:

- After examining the mutual correlations among these parameters, a PCA of the MEGA dataset shows that the 3D parameter space can be described by a plane, dubbed the FPZ.
- The functional form of the (PP04N2) FPZ is given by $12 + \log(O/H) = -0.14 \log(SFR) + 0.37 \log(M_*) + 4.82$ over the entire mass and redshift range of the MEGA dataset.
- The mean O/H residuals of the FPZ over the MEGA dataset are $0.16$ dex for the (PP04N2) calibration, slightly larger for D02 and PP04O3N2; such residuals are smaller than those found previously, and consistent with trends found in smaller galaxy samples with more limited ranges in $M_*$, SFR, and O/H.
- The FPZ is also found to be roughly invariant with redshift enabling an estimation of metallicity accurate to within $0.16$ dex over roughly 5 orders of magnitude in $M_*$, from $10^6 M_\odot$ to $10^{11} M_\odot$, up to $z \approx 3.7$. An additional correction for redshift may be employed to increase slightly the accuracy of O/H estimates from the FPZ for $z \gtrsim 2$ (see Eqn. 5).

ACKNOWLEDGMENTS

We acknowledge the anonymous referee whose insightful comments greatly improved the paper. We also thank G. Cresci for passing us the SDSS10 (KD02) data in electronic form, and are grateful to the DAVID network (http://wiki.arcetri.astro.it/DAVID/WebHome) for fostering a fruitful collaborative environment. PD gladly acknowledges funding from the EU COFUND Rosalind Franklin program.

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