Asymptotic Sum-Capacity of Random Gaussian Interference Networks Using Interference Alignment

Matthew Aldridge Oliver Johnson
Department of Mathematics
University of Bristol, UK
{m.aldridge, o.johnson}@bristol.ac.uk

Robert Piechocki
Centre for Communications Research
University of Bristol, UK
r.j.piechocki@bristol.ac.uk

Abstract—We consider a dense \( n \)-user Gaussian interference network formed by paired transmitters and receivers placed independently at random in Euclidean space. Under natural conditions on the node position distributions and signal attenuation, we prove convergence in probability of the average per-user capacity \( C_S/n \) to \( \frac{1}{2} \mathbb{E} \log(1 + 2 \text{SNR}) \).

The achievability result follows directly from results based on an interference alignment scheme presented in recent work of Nazer et al. Our main contribution comes through the converse result, motivated by ideas of ‘bottleneck links’ developed in recent work of Jafar. An information theoretic argument gives a capacity bound on such bottleneck links, and probabilistic counting arguments show there are sufficiently many such links to tightly bound the sum-capacity of the whole network.

Index Terms—Networks, capacity, sum-capacity, interference alignment, interference network.

I. INTRODUCTION

Recently, progress has been made on many-user approximations to the sum-capacity \( C_S \) of random Gaussian interference networks.

In particular, in a 2009 paper, Jafar [5] proved a result on the asymptotic sum-capacity of a particular random Gaussian interference network:

Theorem 1 [5], Theorem 5). Suppose direct SNRs are fixed and identical, so \( \text{SNR}_i = \text{SNR} \) for all \( i \), and suppose that all TNRs are IID random and supported on some neighbourhood of \( \text{SNR} \). Then the average per-user capacity \( C_S/n \) tends in probability to \( \frac{1}{2} \log(1 + 2 \text{SNR}) \) as \( n \to \infty \).

(Here and elsewhere, we use \( C_S \) to denote the sum-capacity of the network, and interpret \( C_S/n \) as the average per-user capacity.)

A subsequent result by the current authors [6] concerned a more physically realistic model:

Theorem 2 [6], Theorem 1.5). Suppose receivers and transmitters are placed IID uniformly at random on the unit square \([0, 1]^2\), and suppose that signal power attenuates like a polynomial in 1/distance. Then the average per-user capacity \( C_S/n \) tends in probability to \( \frac{1}{2} \mathbb{E} \log(1 + 2 \text{SNR}) \) as \( n \to \infty \).

In this paper, we prove a similar – but more general – result to Theorem 2, with a neater proof, using ideas from Jafar’s proof of Theorem 1. We assume transmitters and receivers are situated independently at random in space (not necessarily uniformly), and that the power of signals depends in a natural way on the distance they travel.

Specifically our result is the following (full definitions of non-italicised technical terms are in Section II):

Theorem 3. Consider a Gaussian interference network formed by \( n \) pairs of nodes placed in an spatially-separated IID network with power law attenuation. Then the average per-user capacity \( C_S/n \) converges in probability to \( \frac{1}{2} \mathbb{E} \log(1 + 2 \text{SNR}) \), in that for all \( \epsilon > 0 \)

\[
\mathbb{P} \left( \left| \frac{C_S}{n} - \frac{1}{2} \mathbb{E} \log(1 + 2 \text{SNR}) \right| > \epsilon \right) \to 0 \quad \text{as} \quad n \to \infty.
\]

The direct part of the proof uses interference alignment. Interference alignment is a new way of dealing with interference in networks, particularly when that interference is of a similar strength to the desired signal. Interference alignment allows communication at faster rates than traditional resource division strategies such as time-division or frequency-division multiple-access. Two early papers on interference alignment are those by Maddah-Ali, Motahari and Khandani [8] and Cadambe and Jafar [3].

Specifically, we take advantage of so-called ergodic interference alignment, developed by Nazer, Gastpar, Jafar and Vishwanath [9].

The converse part of the proof uses the idea of ‘bottleneck links’ developed by Jafar [5]. An information theoretic argument gives a capacity bound on such bottleneck links, and probabilistic counting arguments show there are sufficiently many such links to tightly bound the sum-capacity of the whole network.

A different approach towards finding the capacity of large communications networks is given by the deterministic approach of Avestimehr, Diggavi and Tse [1]. This paper shows how capacities can be calculated up to a gap determined by the number of users \( n \), across all values of \( \text{SNR} \). However, we identify a sharp limit as the number of users \( n \) tends to infinity.

A wider literature review is available in our previous paper [6].

The plan of this paper is as follows: In Section II we define our network model. In Section III we prove the direct part of Theorem 3. Our main contribution comes in Section IV where
we use new ideas to prove the converse part of Theorem 3. We conclude in Section V.

II. MODEL

A. Node position model

We believe that our techniques should work in a variety of models for the node positions. We outline one very natural scenario here.

These ideas were introduced in our earlier paper [6], but were not fully exploited, due to paper’s concentration on the uniform case.

Definition. Consider two probability distributions $\mathbb{P}_T$ and $\mathbb{P}_R$ defined on $d$-dimensional space $\mathbb{R}^d$. Given an integer $n$, we sample the $n$ transmitter node positions $T_1, \ldots, T_n$ independently from the distribution $\mathbb{P}_T$. Similarly, we sample the $n$ receiver node $R_1, \ldots, R_n$ positions independently from distribution $\mathbb{P}_R$. We refer to such a model of node placement as an ‘IID network’.

Equivalently, we could state that transmitter and receiver positions are distributed according to two independent (non-homogeneous) Poisson processes, conditioned such that there are $n$ points of each type. We pair the transmitter and receiver nodes up so that transmitter $i$ at $T_i$ wishes to communicate with receiver $i$ at $R_i$ for each $i$. We make the following definition:

Definition. Let $T \sim \mathbb{P}_T$ and $R \sim \mathbb{P}_R$ be placed independently in $\mathbb{R}^d$. We say the IID network is spatially separated if there exists constants $\beta$ and $c_{sep}$ such that for all $\rho$

$$\mathbb{P}(\|T - R\| \leq \rho) \leq c_{sep}\rho^\beta.$$ 

In particular, it can be shown [6] proof of Lemma 2.2 that the standard dense network is spatially separated. The $d$-dimensional standard dense network is defined by $\mathbb{P}_T$ and $\mathbb{P}_R$ being independent uniform measures on $[0,1]^d$. The standard dense network has been the subject of much research (see for example the review paper of Xue and Kumar [11] and references therein). However, we emphasise that our result holds for a wider range of models.

B. Transmission model

Our results are in the context of so-called ‘line of sight’ communication models, without multipath interference. That is, we consider a model where signal strengths attenuate deterministically with distance according to some function $a$.

The definitions in this section are adapted from our previous paper [6].

Definition. Fix transmitter node positions $\{T_1, \ldots, T_n\} \in \mathbb{R}^d$ and receiver node positions $\{R_1, \ldots, R_n\} \in \mathbb{R}^d$, and consider Euclidean distance $\|\|$ and an attenuation function $a$. We define $\|X\|_a = a(\|X\|)$ and $\|X - Y\|_a = a(\|X - Y\|)$. We define $\|X_i\|_a = a(\|X_i\|)$ and $\|X_i - Y_j\|_a = a(\|X_i - Y_j\|)$.

We consider the $n$-user Gaussian interference network defined so that transmitter $i$ sends a message encoded as a string of $T$ complex numbers $x_i = (x_i[1], \ldots, x_i[T])$ to receiver $i$, under a power constraint $\frac{1}{T} \sum_{t=1}^T |x_i[t]|^2 \leq 1$ for each $i$.

Our result requires that the fading random variables be circularly symmetric. For definiteness, we hold the modulus constant and choose the argument uniformly at random. (We discuss Rayleigh fading in Section V.) So the $t$th symbol received at receiver $j$ is given as

$$Y_j[t] = \exp(i\Theta_{ij}[t]) \frac{\text{SNR}_{ij} x_j[t]}{\sum_{i \neq j} \exp(i\Theta_{ij}[t]) \sqrt{\text{INR}_{ij} x_i[t] + Z_j[t]}},$$

where noise terms $Z_j[t]$ are independent standard complex Gaussian random variables, and the phases $\Theta_{ij}[t]$ are independent $U[0, 2\pi]$ random variables independent of all other terms. The $\text{INR}_{ij}$ and $\text{SNR}$ remain fixed over time, since the node positions themselves are fixed, but the phases are fast-fading, in that they are renewed for each $t$.

Definition. We say an attenuation function $a$ has power law attenuation if there exist constants $\alpha$ and $c_{\text{att}}$ such that for all $\rho$

$$a(\rho) \leq c_{\text{att}}\rho^{-\alpha}.$$ 

(Tse and Viswanath [10] Section 2.1] discuss a variety of models under which power law attenuation is an appropriate model for different exponents $\alpha$.)

For brevity, we write $S_{ij}$ for the random variables $\frac{1}{T} \log(1 + 2\text{INR}_{ij})$ (when $i \neq j$), and $S_{ii}$ for $\frac{1}{T} \log(1 + 2\text{SNR}_i)$ which are functions of the distance between the transmitters and receivers. In particular, since the nodes are positioned independently, under this model the random variables $S_{ij}$ are identically distributed, and $S_{ij}$ and $S_{kl}$ are IID when $\{i,j\}$ and $\{k,l\}$ are disjoint.

We will also write $E = \mathbb{E} S_{ii} = \frac{1}{T} \mathbb{E} \log(1 + 2\text{SNR}_i)$, noting that this is independent of $i$. (It is also true that $E = \mathbb{E} S_{ij}$ for all $i$ and $j$.)

III. PROOF: DIRECT PART

We can now prove our main theorem, Theorem 3, by breaking the probability into two terms which we deal with separately. So

$$\mathbb{P} \left( \left| \frac{C_\Sigma}{n} - E \right| > \epsilon \right) = \mathbb{P} \left( \frac{C_\Sigma}{n} - E < -\epsilon \right) + \mathbb{P} \left( \frac{C_\Sigma}{n} - E > \epsilon \right).$$

Bounding the first term of (2) corresponds to the direct part of the proof. Bounding the second term of (2) corresponds to the converse part, and represents our major contribution. We prove the direct part as previously [6].

Proof: The first term of (2) can be bounded relatively simply, using an achievability argument based on an interference alignment scheme presented by Nazer, Gastpar, Jafar and Vishwanath [9]. Their theorem [9, Theorem 3] implies that the rates $R_i = 1/2 \log(1 + 2\text{SNR}_i) = S_{ii}$ are simultaneously...
achievable. This implies that $C_2 \geq \sum_{i=1}^{n} R_i = \sum_{i=1}^{n} S_{ii}$. This allows us to bound the first term in (2) as
\[
\mathbb{P}\left( \frac{C_2}{n} - E < -\epsilon \right) \leq \mathbb{P}\left( \frac{\sum_{i=1}^{n} S_{ii}}{n} < E - \epsilon \right).
\]
But $E = \mathbb{E} S_{ii}$, so this probability tends to 0 by the weak law of large numbers.

IV. PROOF: CONVERSE PART

We now need to show that the second term of (4) tends to 0 too. Specifically, we must prove the following: for all $\epsilon > 0$
\[
\mathbb{P}\left( \frac{C_2}{n} \geq E + \epsilon \right) \to 0 \quad (3)
\]
as $n \to \infty$.

The proof of the converse part is the major new part of this paper. First, bottleneck links are introduced, and we prove a tight information-theoretic bound on the capacity of such links. Second, a probabilistic counting argument ensures there are (with high probability) sufficiently many bottleneck links to bound the sum-capacity of the entire network.

A. Bottleneck links

The important concept is that of the bottleneck link, an idea first used by Jafar [5] and later adapted [6] in the following form:

**Definition.** We say the link $i \rightarrow j$, $i \neq j$, is an $\epsilon$-bottleneck link, if the the following three conditions hold:

B1: $S_{ii} \leq E + \epsilon / 2$,

B2: $S_{ij} \leq E + \epsilon / 2$,

B3: $S_{jj} \leq S_{ji}$.

We let $B_{ij}$ be the indicator function that the crosslink $i \rightarrow j$ is a $\epsilon$-bottleneck link. We also define the bottleneck probability $\beta := E B_{ij}$ to be the probability that a given link is an $\epsilon$-bottleneck which is independent of $i$ and $j$ for an IID network. (We suppress the $\epsilon$ dependence for simplicity.)

The crucial point about bottleneck links is the constraints they place on achievable rates in a network.

**Lemma 4.** Consider a crosslink $i \rightarrow j$ in a $n$-user Gaussian interference network. If $i \rightarrow j$ is a $\epsilon$-bottleneck link, then the sum of their achievable transmission rates is bounded by $r_i + r_j \leq 2E + \epsilon$.

**Proof:** First, note that we make things more nearly by considering the two-user subnetwork:

\[Y_i = \exp(i\Theta_{ii})\sqrt{\text{SNR}_i}X_i + \exp(i\Theta_{ij})\sqrt{\text{INR}_{ij}}X_j + Z_i\]
\[Y_j = \exp(i\Theta_{ij})\sqrt{\text{INR}_{ij}}X_i + \exp(i\Theta_{jj})\sqrt{\text{SNR}_j}X_j + Z_j\]

where receiver $i$ needs to determine signal $X_i$, and receiver $j$ signal $X_j$. (The time index is omitted for clarity.)

From bottleneck conditions B1 and B2 we have
\[1 + 2\text{SNR}_i \leq \exp(2E + \epsilon), \quad 1 + 2\text{INR}_{ij} \leq \exp(2E + \epsilon).
\]

Summing and taking logs gives
\[\log(1 + \text{SNR}_i + \text{INR}_{ij}) \leq 2E + \epsilon. \quad (4)\]

We combine this with the argument given by Jafar [5]. Let $r_i$ and $r_j$ be jointly achievable rates for the subnetwork. In particular, receiver $i$ can determine signal $X_i$ with an arbitrarily low probability of error.

We certainly do no worse if a genie presents signal $X_i$ to receiver $j$—so assume $j$ can indeed recover $X_i$. But condition B3 ensures that it is easier for receiver $i$ to determine $X_j$ than it is for receiver $j$ (since the weighting is larger in the first case). So since receiver $j$ can recover $X_j$ (as $r_j$ is achievable), receiver $i$ can recover $X_j$ also.

Because receiver $i$ can determine both $X_i$ and $X_j$, these two signals must have been transmitted at a sum-rate no higher than the sum-capacity of the Gaussian multiple-access channel (see, for example, Cover and Thomas [4] Section 14.3.6). Hence,
\[r_i + r_j \leq \log(1 + \text{SNR}_i + \text{INR}_{ij}) \leq 2E + \epsilon,\]
where the second inequality comes from (4).

B. Three technical lemmas

A few technical lemmas are required in order to prove (3).

First, we need to ensure that very high SNRs are very rare (Lemma 5). Second, we need to show that bottleneck links will actually occur (Lemma 6). Last, we must show that the number of bottleneck links cannot vary too much (Lemma 7).

Under any network model where these three lemmas are true, our theorem will hold. We emphasise that our model of IID networks with power law attenuation is one such model; we believe the result holds more widely.

**Lemma 5.** Consider a spatially-separated IID network, with power law attenuation. Then for any $\eta > 0$,
\[\mathbb{P}\left( \max_{1 \leq i \leq n} S_{ii} > n^{\eta/2} \right) = O(n^{-1}) \quad \text{as } n \to \infty.
\]

In fact, in our case the convergence to 0 is considerably quicker than $O(n^{-1})$, but this is sufficient.

It is worth noting that this fast decay in the tails of $S_{ii}$ ensures that the expectation $E = \mathbb{E} S_{ii}$ does indeed exist and is finite.

**Proof:** First, we have by the union bound
\[\mathbb{P}(\max_{1 \leq i \leq n} S_{ii} > n^{\eta/2}) \leq n \mathbb{P}(S_{11} > n^{\eta/2}).\]

But by the definition of $S_{11}$
\[\mathbb{P}(S_{11} > n^{\eta/2}) = \mathbb{P}\left( \frac{\text{SNR}_{11}}{2} > \frac{1}{2}(2n^{\eta/2} - 1) \right) = \mathbb{P}\left( a(||\mathbf{T}_1 - \mathbf{R}_1||) > \frac{1}{2}(2n^{\eta/2} - 1) \right)
\]
and the proof follows by applying the definitions of SNR, spatial separation and power law attenuation.

We will often condition on this event: that is, condition on the complementary event $\{\max_{1 \leq i \leq n} S_{ii} \leq n^{\eta/2}\}$. We use $\mathbb{P}_n$, $\mathbb{E}_n$ and $\text{Var}_n$ to denote such conditionality, and write $\beta_n = \mathbb{E}_n B_{ij}$ for the conditional bottleneck probability.
The next two lemmas concern showing that conditional probabilities are nonzero. However, we have for any event $A$,
$$\Pr(A) = \Pr(A \mid \max S_{ii} \leq n^{n/2}) \Pr(\max S_{ii} \leq n^{n/2})$$
$$+ \Pr(A \mid \max S_{ii} > n^{n/2}) \Pr(\max S_{ii} > n^{n/2}).$$
and hence by Lemma 5 we have the bounds
$$\Pr(A) \leq \Pr(A \mid \max S_{ii} \leq n^{n/2}) + \Pr(\max S_{ii} > n^{n/2})$$
$$= \Pr_n(A) + O(n^{-1})$$
and
$$\Pr(A) \geq \Pr(A \mid \max S_{ii} \leq n^{n/2}) \Pr(\max S_{ii} \leq n^{n/2})$$
$$= \Pr_n(A)(1 - O(n^{-1})),$$
and so $\Pr(A) = \Pr_n(A) + O(n^{-1}).$ This will be useful in the next two proofs.

**Lemma 6.** Consider a spatially-separated IID network, with power law attenuation. Then the conditional bottleneck probability $\beta_n$ is bounded away from 0 for all $n$ sufficiently large.

**Proof:** First note that by the comment above, we need only show that the unconditional bottleneck probability $\beta$ is nonzero.

Second, note that by the exchangeability of $R_i$ and $R_j$, we have $\Pr(B1 \text{ and } B2 \text{ and } B3) \geq \frac{1}{2} \Pr(B1 \text{ and } B2)$. It is left to show that $\Pr(B1 \text{ and } B2)$ is non-zero.

Note that $B1$ requires $S_{ii}$ to be less than its expectation plus $c$. So $T_i$ must be situated such that this has nonzero probability. So $T_i$ has a nonzero probability of being positioned such that $B1$ occurs. But $T_i$ and $T_j$ are also exchangeable, so we are done.

**Lemma 7.** Consider a spatially-separated IID network, with power law attenuation. Then, conditional on $\{\max S_{ii} < n^{n/2}\}$,
$$\Var_n(\text{# bottleneck links}) = \Var_n\left(\sum_{i \neq j} B_{ij}\right) = O(n^3),$$
where the sum is over all crosslink pairs $(i, j), i \neq j$.

In general, one might assume that $\Var_n(\text{# bottleneck links})$ would be proportional to the total number of links, and thus be $O(n^3).$ However, because of the independences in the IID network, the variance is in fact much lower.

**Proof:** First consider the unconditional version. We have
$$\Var\left(\sum_{i \neq j} B_{ij}\right) = \sum_{i \neq j} \sum_{k \neq l} \Cov(B_{ij}, B_{kl}).$$

The important observation is that for $i, j, k, l$ all distinct, $B_{ij}$ and $B_{kl}$ are independent giving $\Cov(B_{ij}, B_{kl}) = 0.$ (This is because they depend only on the position of distinct and independently-positioned nodes.) Hence there are only $O(n^3)$ non-zero terms in the sum, each of which is trivially bounded by $\frac{1}{4}(1 - \frac{1}{4}) = \frac{1}{4}$.

But by the comment above, if $\Cov(B_{ij}, B_{kl}) = 0$, then the conditional covariance is $\Cov_n(B_{ij}, B_{kl}) = O(n^{-1})$. Hence,
$$\Var_n\left(\sum_{i \neq j} B_{ij}\right) \leq O(n^3)\frac{1}{4} + O(n^4)O(n^{-1}) = O(n^3),$$
as desired.

**C. Completing the proof of Theorem 3**

We are now in a position to prove 4, and hence prove Theorem 3.

**Proof:** We need to show
$$\forall \epsilon > 0 \forall \delta > 0 \exists N \forall n \geq N \Pr\left(\frac{C\cdot \log n}{n} \geq E + \epsilon\right) \leq \delta.$$
So choose $\epsilon > 0$, $\delta > 0$, and fix $n \geq N$ (where $N$ will be determined later), and pick a rate vector $r \in \mathbb{R}_+^n$ with sum-rate
$$\frac{r\cdot \log n}{n} > E + \epsilon;$$
for $N$ sufficiently large, by Lemma 5 we need to bound the first term in 8.

First, note that our assumption on $\max S_{ii}$ means that if $r_i > 2n^{n/2}$, then we break the single-user capacity bound, since we would have
$$r_i > 2n^{n/2} \geq 2 \max S_{ij}$$
$$\geq 2S_{ii} = \log(1 + 2\operatorname{SNR}_i) > \log(1 + \operatorname{SNR}_i)$$
meaning $r$ is not achievable, and we are done. Thus we assume this does not hold; that
$$r_i \leq 2n^{n/2} \text{ for all } i.$$ (7)
(The rest of our argument closely follows Jafar 5.)

Now, if $r$ is achievable, it must at least satisfy the constraints on the $\epsilon$-bottleneck links $i \rightarrow j$ from Lemma 4 and hence also the sum of those constraints. So
$$\Pr_n(r \text{ achievable})$$
$$\leq \Pr_n(r_i + r_j \leq 2E + \epsilon \text{ on bottleneck links } i \rightarrow j)$$
$$\leq \Pr_n\left(\sum_{i \neq j} B_{ij}(r_i + r_j) \leq \left(\sum_{i \neq j} B_{ij}\right)(2E + \epsilon)\right)$$
$$= \Pr_n(U \leq V),$$ (8)
where we have defined
\[ U := \frac{1}{n(n-1)} \sum_{i \neq j} B_{ij}(r_i + r_j), \]
\[ V := \frac{1}{n(n-1)} \left( \sum_{i \neq j} B_{ij} \right) (2E + \epsilon). \]

The conditional expectations of \( U \) and \( V \) are
\[ \mathbb{E}_n U = 2\beta_n \frac{\gamma_n}{n}, \quad \mathbb{E}_n V = \beta_n (2E + \epsilon) = 2\beta_n \left( E + \frac{\epsilon}{2} \right). \]

Note that since \( \beta_n > 0 \) by Lemma 6, we can rewrite (5) as \( \mathbb{E}_n U > \mathbb{E}_n V + \beta_n \epsilon \), or equivalently,
\[ \mathbb{E}_n U - \frac{\beta_n \epsilon}{2} > \mathbb{E}_n V + \frac{\beta_n \epsilon}{2}. \]

The proof is completed by formalising the following idea: since the expectations are ordered \( \mathbb{E}_n U > \mathbb{E}_n V \), we can only rarely have the opposite ordering \( U < V \). Hence the expression in (8) is small.

Formally, by (the conditional version of) Chebyshev’s inequality and the union bound, we have
\[ \mathbb{P}_n(U \leq V) \]
\[ \leq \mathbb{P}_n \left( U \leq \mathbb{E}_n U - \frac{\beta_n \epsilon}{2} \text{ or } V' \geq \mathbb{E}_n V + \frac{\beta_n \epsilon}{2} \right) \]
\[ \leq \mathbb{P}_n \left( |U - \mathbb{E}_n U| \geq \frac{\beta_n \epsilon}{2} \right) + \mathbb{P}_n \left( |V - \mathbb{E}_n V| \geq \frac{\beta_n \epsilon}{2} \right) \]
\[ \leq \frac{1}{\beta_n \epsilon} \text{Var}_n U + \left( \frac{2}{\beta_n \epsilon} \right)^2 \text{Var}_n V \]
\[ = \frac{4}{\beta_n^2 \epsilon^2} (\text{Var}_n U + \text{Var}_n V). \tag{9} \]

Using Lemma 7 we can bound these variances as
\[ \text{Var}_n U = \frac{1}{n^2(n-1)^2} \text{Var}_n \left( \sum_{i \neq j} B_{ij}(r_i + r_j) \right) \]
\[ \leq \frac{1}{n^2(n-1)^2} O(n^3) 16n^n = O(n^{-(1-n)}), \]
\[ \text{Var}_n V = \frac{1}{n^2(n-1)^2} O(n^3)(2E + \epsilon)^2 = O(n^{-1}), \]

where we used (2) to bound \( \text{Var}_n U \). Choosing \( \eta \) to be less than 1, we can ensure \( N \) is sufficiently large that for all \( n \geq N \)
\[ \text{Var}_n U + \text{Var}_n V \leq \frac{\beta_n^2 \epsilon^2}{8}. \]

This makes \( \mathbb{P}_n(U \leq V) < \delta/2. \) Together with (8) and (9), this yields the result.

V. CONCLUSION

In this paper we have defined IID interference networks with power law attenuation. We have shown that this setup fulfills necessary properties for the average per-user capacity \( C_{\Sigma}/n \) to tend in probability to \( \frac{1}{2} \mathbb{E} \log(1 + 2\text{SNR}) \). We have also noted that this result is not unique to our setup.

We briefly mention one more example. Suppose Rayleigh fading is added to our model. That is, now let \( \text{SNR}_i := |H_{ii}|^2 \text{cov}(|T_i - R_{ij}|) \) and \( \text{INR}_{ij} := |H_{ij}|^2 \text{cov}(|T_i - R_{ij}|) \), where the \( H_{ij} \) are IID standard complex Gaussian random variables. Because ergodic interference alignment still works with Rayleigh fading [9], the direct part of the theorem still holds. But also, because the fading coefficients are IID, the independence structure from the non-fading case remains, ensuring Lemmas 5–7 hold. Hence, the theorem is still true.

Characterising all networks for which such a limit for average per-user capacity exists is an open problem.

At the moment, Theorem 3 should perhaps be regarded as being of theoretical interest. That is, our major contribution is to provide a sharp upper bound on the performance of interference networks. However, the lower bound relies on an ergodic interference alignment [9] which, while rigorously proved, may not be feasible to implement in practice for large number of users. Examination of the proof of the effectiveness of ergodic interference alignment [9] Theorem 1] shows that, even for a model with alphabet size \( q \), the channel needs to be used \( O((q - 1)^K) \) times. Even for \( K \approx 10 \), this is a prohibitive requirement. However, recent work by the current authors [7] characterises the delay–rate tradeoff for ergodic interference alignment. Also, note that for \( K = 3 \), El Ayach, Peters, and Heath [2] have shown that the interference alignment scheme of Cadambe and Jafar [3] can perform close to the theoretical bounds.

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