Influence of spatially variable material properties on the resistance of masonry walls under compression

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Abstract During their assessment, existing masonry structures often reveal a high variability of their material properties. Since this variability is also present within a single wall, the influence of this spatial variability of the material properties on the wall resistance is of major interest. First, experimental investigations on clay brick masonry walls with local weaknesses under compression loading are presented. The results of these experiments serve to validate a finite element model developed to study the influence of spatially variable material properties on the resistance of masonry walls under compression. The model follows the simplified micro-modelling approach and allows for a unit-to-unit variability of the compressive strength and the elastic modulus of masonry. The finite element model is then utilised within Monte Carlo simulations to investigate the resulting probability distribution of the wall resistance for given distributions of the material properties. Finally, parameter studies are conducted to examine the influence of the coefficient of variation of the material properties, the spatial correlation of the material properties, wall length, and wall slenderness on the resulting distribution of the resistance. The studies demonstrate that spatial variability leads to a decrease in the mean value of the wall resistance. However, the resulting variability of the resistance is much smaller than the input variability of the material properties. Considering the spatial nature of the variability can thus lead to higher design/assessment values of the wall resistance and, hence, a positive effect on structural reliability.

Keywords Masonry · Monte Carlo simulation · Finite element modelling · Existing structures · Probabilistic assessment

1 Introduction

1.1 Motivation

The assessment of existing masonry structures involves various types of uncertainty, making it a challenging task [1]. Since the number of destructive material tests is usually kept to a minimum due to economic and structural reasons and, in some cases, to preserve cultural heritage, statistical uncertainty must be considered; the low number of test results might not be representative of the whole population of masonry. Furthermore, a high level of model uncertainty must be taken into account. In addition to the uncertainty of the resistance model, masonry compressive strength can also be associated with considerable model
uncertainty, as masonry compressive strength is usually determined by separate tests on unit and mortar specimens and subsequent application of an empirical model. Approaches to consider statistical and model uncertainty in the assessment of existing masonry structures have been presented in [2].

Apart from statistical and model uncertainty, there is spatial variability of the material properties: Material properties do not only vary from wall to wall but also within a wall. This spatial variability is also present in new masonry structures but becomes particularly apparent in the assessment of older masonry structures with high material variability. The effects of spatial material variability on the out-of-plane flexural resistance and the in-plane shear resistance of masonry walls have been investigated in [3–5] and [6], respectively, through Monte Carlo simulations (MCS) employing finite element models. According to these investigations, considering spatial variability causes a decrease in the mean resistance compared to a non-spatial analysis with unit and mortar properties modelled as homogeneous within a masonry wall. However, if spatial variability is considered, the coefficients of variation (CoVs) of the wall resistances are much smaller than the input CoVs of the most influential material properties.

The influence of spatially variable material properties on the resistance of masonry under compression has already been examined by Goretzky [7]. Goretzky’s goal was to compare the resistance of masonry with high spatial variability of strength and elastic modulus of unit and mortar, which is typical for historic masonry, to that of modern masonry with lower variability. For reducing the computational effort associated with using a finite element model, Goretzky developed a simplified numerical model representing masonry as a set of vertical prisms. MCS were then conducted to derive reduction factors for the mean value and the characteristic value (5% fractile) of the resistance in relation to that of contemporary masonry. Lower fractile values, which are essential concerning structural reliability, were not evaluated in [7]. The simulations in [7] were only carried out for a masonry wall with a length of six units per course and a height of only seven courses. Only concentric compression loading and no second-order effects were considered. Therefore, further investigations with varying boundary conditions, such as wall length and slenderness, are needed for a deeper understanding of the influence of spatial variability on masonry under compression.

Current safety concepts for the design and assessment of structures usually do not take spatial variability into account. Instead, masonry compressive strength, for example, is assumed to show only variability from wall to wall. Within a wall, uniformity of the compressive strength (i.e. homogeneity) is assumed. The design value of masonry compressive strength can then be directly determined as a lower quantile value of the respective probability distribution according to the approach with fixed sensitivity factors specified in Eurocode 0 [8]. In the assessment of existing structures, the variance of this probability distribution can be based on (or updated with) results from material tests. For a reliable assessment of existing, as well as a reliable design of new masonry, it is essential to know how these design values differ if they are determined under consideration of spatial variability.

1.2 Objective and outline

To illustrate the objective of the following study, the general effects of considering spatial variability are presented in Fig. 1. As an example, the figure considers a masonry wall under concentric compression that is not subjected to second-order effects (i.e. with theoretical slenderness of zero). In this case, the resistance is mainly determined by masonry compressive strength. If normalised by cross-section $A$ and mean compressive strength $f_{ma,m}$, the probability density function (PDF) of the resistance $R_{hom}$ of a homogeneous wall (i.e. the resistance based on a non-
spatial analysis) equals the PDF of masonry compressive strength \( f_{ma} \) normalised by its mean \( f_{ma,m} \). The means of the normalised resistance \( R_{nom} \) and the normalised masonry compressive strength \( f_{ma} \) are one, equaling the normalised resistance \( R_{det} \) based on a deterministic calculation with \( f_{ma,m} \) as input material property. If spatial variability is considered, the two previously described effects set in: both the mean and the CoV reduce compared to a non-spatial analysis. Furthermore, the probability distribution of the resistance does not follow the same distribution type as the input material property (here: log-normal distribution for masonry compressive strength).

Concerning design values for masonry compressive strength, as well as the reliability of masonry walls in general, the left tail of the PDF is essential. To determine the overall effect of spatial variability, the relationship between the probability distribution of spatially variable masonry compressive strength and the resulting probability distribution of the resistance of a masonry wall under compression must be known. Therefore, two main questions need to be answered:

1. How does spatial variability influence the mean value of the resistance of a masonry wall in comparison to that of a homogeneous masonry wall?
2. How can the relationship between the variability of material properties and the resulting variability of the wall resistance be described?

If these two effects are known, and a suitable distribution type is selected, design values for masonry compressive strength can be determined according to the current Eurocode safety format. In the assessment of existing structures, design values are sometimes referred to as assessment values to highlight the different perspectives (see e.g. [9]). Here, the term “design value” is kept to match the well-known nomenclature of the current Eurocodes.

The following analyses focus on solid clay brick masonry walls under compression. In Sect. 2, experimental investigations to study the influence of local weaknesses on the resistance of solid clay brick masonry walls (i.e. their stress redistribution capability) are presented. The experimental results serve as a basis for validating a finite element model, which is illustrated in Sect. 3. The finite element model is then used to conduct Monte Carlo simulations (MCS), in which the compressive strength, elastic modulus, and flexural tensile strength of masonry are modelled spatially variable. The corresponding stochastic model is described in Sect. 4. Finally, the general approach of the MCS and the results of parameter studies are presented in Sect. 5.

2 Experimental investigations

2.1 Experimental programme

The experiments aimed to investigate the behaviour of solid clay brick masonry with local weaknesses. In total, 24 masonry walls were tested, which were either arranged as single wythe masonry or in cross bond (see Fig. 2). In addition to reference walls, walls with intentionally placed weaknesses were constructed. These weaknesses consisted of either a missing brick in the masonry bond or a particular percentage (25% or 50%) of perforated clay bricks (i.e. bricks with significantly lower compressive strength) within the walls. In walls with a missing brick, stress concentrations occur close to the hole. When the masonry strength is reached locally, stresses need to be redistributed to areas with lower stress levels. The resistance of such walls is thus mainly determined by the post-peak behaviour of the material, making these walls well suited for the validation of the finite element model. In reality, walls with high spatial variability do not only feature a single weak spot but various weaknesses that are randomly distributed over the wall geometry. Furthermore, the weaknesses consist of weaker bricks (i.e. they are less extreme than a hole). Hence, an additional mechanical influence is the confinement of weaker bricks by stronger ones. Therefore, the walls including both stronger and weaker bricks complemented the experimental programme by walls with more realistic boundary conditions to capture the major influences in the validation of the finite element model.

All the tested walls had a height of 13 courses with five whole bricks per course. For mortar, a factory-mixed dry mortar containing natural hydraulic lime NHL 5 according to EN 459-1 [10], pozzolans, and sand with a maximum aggregate size of 1.2 mm, was chosen. The nominal perpend and bed joint thicknesses were selected as 10 mm and 12 1/3 mm, respectively. The nominal brick dimensions were \( 240 \times 115 \times 71 \) mm\(^3\) (length \( \times \) width \( \times \) height),
corresponding to standard format NF according to DIN 20000-401 [11]. Consequently, the nominal dimensions of the single wythe walls were 1240 × 115 × 1083 mm$^3$ and those of the cross bonded walls 615 × 115 × 1083 mm$^3$.

The walls were erected by trained masons and covered by polyethylene sheets for the first 3 days after construction. Then, until the testing at the age of 32 up to 42 days, the walls were stored under laboratory conditions. Before testing, the top of the wall was capped with a thin gypsum layer. A concentric compression load was then applied via a steel beam, whose support allowed for rotation, with a constant displacement rate to reach the maximum load after 15 to 30 min. Displacements were measured by four linear variable displacement transducers (LVDTs) per wall, each reaching over 10 brick and bed joints. The test setup is depicted in Fig. 3.

In addition to the tests on masonry walls, the results of which are presented in Sect. 2.3, complementary compression tests on bricks, mortar specimens, and masonry specimens were performed, as presented in Sect. 2.2.

2.2 Results of tests on unit, mortar, and masonry specimens

Unit compression tests according to EN 772-1 [12] were performed on at least $n = 6$ bricks from each delivered pallet of bricks. The results for the solid and perforated clay bricks are listed in Table 1. In addition to the directly obtained strengths, the normalised
compressive strengths based on the shape factors defined in [12] are given. The average normalised unit compressive strengths were 24.9 and 11.6 N/mm² for solid and perforated bricks, respectively.

For the construction of each wall, several mortar mixes were needed. From at least one of these mixes, three standard mortar prisms were sampled and tested according to EN 1015-11 [13]. The standard prism mortar tests were conducted on the same day as the corresponding tests on masonry walls. On average, a compressive strength of 2.71 N/mm² was obtained on the standard mortar prisms. In addition to the standard prisms tests, double punch tests according to procedure III of DIN 18555-9 [14] were performed on small specimens (50 × 50 × 12 mm³) that were extracted from mortar joints within masonry (see Fig. 4). According to this testing method, the load is applied via two loading platens with a diameter of 20 mm. Since these mortar specimens cured under different conditions, namely within masonry instead of in steel moulds, the obtained strengths are much different: 6.50 and 7.44 N/mm² were determined for mortar cured in solid and perforated clay brick masonry, respectively. It is noted that all the double punch tests were conducted on mortar specimens belonging to the same mortar mix; the average standard prism strength of this mix was 3.48 N/mm² (i.e. slightly higher than the overall average of 2.71 N/mm²). Hence, the double punch strength was higher by a factor of about two, which approximately matches the findings in [15]. For further information on double punch testing, refer to [16–19].

Furthermore, the compressive strength and the modulus of elasticity of masonry were determined according to EN 1052-1 [20] on six masonry specimens (so-called “RILEM specimens”; see [21]) for both masonry types. These masonry specimens, which were arranged as single wythe masonry, had a height of six and a length of two units. The corresponding results are also listed in Table 1.

![Test setup for wall cb-50-1](Fig. 3)

### Table 1  Experimentally obtained material properties

| Specimen type | Property                              | Testing standard | Solid clay bricks | Perforated clay bricks |
|---------------|---------------------------------------|------------------|-------------------|------------------------|
| Unit          | Compressive strength⁴                 | EN 772-1         | 29.7 (0.08)       | 13.8 (0.11)            |
|               | Normalised compressive strength⁴      |                  | 24.9 (0.08)       | 11.6 (0.11)            |
| Mortar        | Standard prism strength⁵              | EN 1015-11       | 2.71 (0.17)       |                        |
|               | Double punch strength⁶                | DIN 18555-9      | 6.50 (0.03)       | 7.44 (0.03)            |
| Masonry       | Compressive strength⁷                 | EN 1052-1        | 9.74 (0.11)       | 4.11 (0.09)            |
|               | Modulus of elasticity⁸                |                  | 3000 (0.20)       | 1959 (0.28)            |

All quantities are average values displayed in N/mm² with CoVs in parentheses

|   |   |   |   |
|---|---|---|---|
|   |   |   |   |

[^4]: n = 37 tests on solid bricks and n = 18 tests on perforated bricks
[^5]: n = 226 tests on mortar prisms
[^6]: n = 12 double punch tests for each masonry type. The average standard prism strength for the corresponding mortar mix is 3.48 N/mm² (n = 12)
[^7]: n = 6 tests for each masonry type
2.3 Results of tests on masonry walls

The resulting load-bearing capacities of the tested masonry walls are displayed in Table 2. In addition to the maximum applied load, the testing age, the wall dimensions, and the strength (maximum load per gross cross-sectional area) are listed. Furthermore, the average strength is given and a relative strength, which is the strength related to the respective solid clay brick reference wall.

The walls with a missing brick (sw-hole, cb-str, cb-head) showed a relative strength of between 70 and 78%. If perfect stress redistribution were possible, the relative load-bearing capacity would be 80%, which is the relative area of the critical cross-section (four instead of five bricks). The strength of walls containing solid and perforated bricks is lower than the strength that is obtained by linear interpolation between the strengths of the two reference walls according to the respective share of solid and perforated bricks. Thus, the load-bearing capacity cannot be directly determined via the average masonry compressive strength in the wall. If only the critical cross-section (i.e. the cross-section with the largest share of perforated bricks) is considered, the extent of this effect reduces, but in principle, the effect remains. Hence, the tests revealed a considerable but not perfect capability of stress redistribution within the walls.

### 3 Finite element modelling

#### 3.1 General model description

The software DIANA FEA (Version 10.3) is utilised for modelling the masonry walls. The model follows the simplified micro-modelling approach [22]. Hence, the walls are represented by expanded units (i.e. units with their dimensions increased by half of the adjacent mortar joint thicknesses) and interfaces with zero thickness, representing the mortar joints. The simplified micro-modelling approach is chosen as it provides a good compromise between the detailed micro-modelling approach, which represents the interaction of unit and mortar in detail, and the macro-modelling approach, where no discretisation between units and mortar joints is present. On the one hand, detailed micro-modelling would be computationally more demanding and would require much more mechanical and stochastic input parameters. On the other hand, the missing discretisation of a masonry wall into separate units makes macro-modelling inapplicable for walls with unit-to-unit material variability.

Here, the expanded units are modelled by eight-node solid elements, to which the nonlinear behaviour of the composite material masonry is assigned. The dimensions of the expanded units are $250 \times 115 \times 83.3$ mm$^3$ (length $\times$ width $\times$ height) for the single wythe masonry walls, corresponding to the nominal dimensions of brick format NF plus two halves of the perpend joint thickness (10 mm) in length and two halves of the bed joint thickness.
(12.3 mm) in height. In the case of the cross-bonded walls, the dimensions are 250 × 120 × 83.3 mm³ for stretchers and 240 × 125 × 83.3 mm³ for headers to obtain a wall width of 240 mm. Each unit is discretised into 8 × 4 × 3 elements. Thus, the elements are approximately cubic with an edge length of approximately 30 mm. The interfaces that are placed at the position of the perpend and bed joints consist of plane quadrilateral elements with four plus four nodes. The purpose of the interfaces is to display the discrete cracking of masonry at the mortar joints under tension.

The load is applied as a line load with a specified eccentricity that is then distributed to the whole cross-section via a rigid plate at the top. At the bottom, the walls are supported vertically along a line, which is similarly connected to a load distribution plate. Furthermore, horizontal supports are placed at the top and bottom. In the simulation of the experiments, the load is applied displacement-controlled, and the top is hinged around the longitudinal axis (x-axis; see Fig. 5 Geometry, mesh, and example strains of the finite element model (reference wall for the Monte Carlo simulations).
Fig. 5). The rotation is restraint at the bottom to model the experimental conditions. The MCS are performed load-controlled. The support at the top allows for rotation around the x- and the y-axis, and the support at the bottom for rotation around the x-axis. Figure 5 displays the geometry, mesh, example stresses, and example strains of the finite element model. The depicted wall corresponds to the reference wall in the MCS.

3.2 Material modelling

Although the material behaviour of masonry is orthotropic in general, the material behaviour of the extended units is modelled as isotropic here. An isotropic model yields sufficiently accurate results for the following investigations since the vertical compressive stresses dominate. The compressive behaviour of the expanded units is modelled with a Drucker–Prager yield criterion combined with a linear tension cut-off and cohesion hardening/softening. The flow rule is chosen to be associated; that is, the plastic compressive strain $\varepsilon_{pl}$, the hardening/softening law is defined as follows:

$$
\sigma_c(\varepsilon_{pl}) = \sigma_i + (\sigma_p - \sigma_i) \left[ 2 \frac{\varepsilon_{pl}}{\varepsilon_p} - \frac{\varepsilon_{pl}^2}{\varepsilon_p^2} \right] \text{ for } \varepsilon_{pl} \leq \varepsilon_p, \sigma_c(\varepsilon_{pl}) = \sigma_p + (\sigma_m - \sigma_p) \left( \frac{\varepsilon_{pl} - \varepsilon_p}{\varepsilon_m - \varepsilon_p} \right)^2 \text{ for } \varepsilon_p \leq \varepsilon_{pl} \leq \varepsilon_m, \sigma_c(\varepsilon_{pl}) = \sigma_i + (\sigma_m - \sigma_i) \times \exp \left( 2 \frac{\sigma_m - \sigma_p \varepsilon_{pl} - \varepsilon_m}{\varepsilon_m - \varepsilon_p \sigma_m - \sigma_i} \right) \text{ for } \varepsilon_m \leq \varepsilon_{pl},
$$

With $\sigma_p$ equal to masonry compressive strength $f_{ma}$ and $\varepsilon_p$ the corresponding plastic strain. For $\sigma_c \leq \sigma_i$, only elastic strains are present. The post-peak behaviour is defined by $\sigma_m$ and $\varepsilon_m$. The residual stress $\sigma_r$ is needed to avoid numerical issues. In [22], the values $\sigma_i = 0.33 f_{ma}$, $\sigma_m = 0.5 f_{ma}$, and $\sigma_r = 0.1 f_{ma}$ are proposed, which are also applied here. Setting $\varepsilon_m = 6 \varepsilon_p$ yields results matching the experiments well.

For a given modulus of elasticity and a given masonry compressive strength, the plastic strain $\varepsilon_p$ is determined by the stress–strain parameter $k$, which is the ratio of total to elastic strain at peak stress. The parameter $k$ is a measure for the nonlinearity of the stress–strain curve, with $k = 1$ leading to a linear curve and $k = 2$ to a curve resembling a quadratic parabola. From the results of tests on the reference walls, $k$ is determined as 2.31 (single wythe) and 2.38 (cross bond) for solid clay brick masonry, and 1.69 (single wythe) and 2.04 (cross bond) for perforated clay brick masonry. In the MCS, $k = 2$ is selected, which is representative of solid clay brick masonry in general [24, 25]. For the MCS, an average ratio $E_{ma}/f_{ma} = 550$ is selected based on [24, 26].

To define the tension cut-off of the material model for the expanded units, the horizontal tensile strength $f_{bs}$ of the units is assigned. For this purpose, $f_{bs}$ is set to 0.04 $f_b$ for the solid bricks and 0.03 $f_b$ for the perforated bricks based on [26], where $f_b$ is the unit compressive strength. The MCS are performed based on material properties that are normalised regarding masonry compressive strength $f_{ma}$. Based on typical unit and mortar strengths of $f_b = 25$ N/mm$^2$ and $f_j = 5$ N/mm$^2$, respectively, a characteristic masonry strength of

![Figure 6](image-url)
\( f_{ma,k} = 8.1 \text{ N/mm}^2 \) is obtained according to [27], corresponding to a mean value of \( f_{ma,m} = 9.1 \text{ N/mm}^2 \) (assuming \( f_{ma,k} = 0.8 \ f_{ma,m} \) [28]). Hence, the normalised unit tensile strength is \( f_{bt}/f_{ma} = 0.11 \). Tension softening is selected as linear with a residual tensile strength of 0.1 \( f_{bt} \).

A discrete cracking material model with bilinear softening according to [29] is assigned to the interfaces that represent the mortar joints. As the deformation behaviour under compression is modelled by the expanded units, the stiffness of the interfaces is set very high such that almost no relative displacements occur if the joint is uncracked. The tensile strength \( f_t \) of the joints represents the flexural tensile strength \( f_{x1} \) of masonry. A typical value of 0.4 N/mm\(^2\) [30] is chosen, corresponding to 0.044 \( f_{ma} \) based on the normalisation described above. The tensile fracture energy is set to 0.0148 \( f_t \) based on findings in [31]. The selected material properties of the finite element model are summarised in Table 3.

### 3.3 Validation of the model

The finite element model needs to display the behaviour of solid clay brick masonry with local weaknesses sufficiently well. In particular, the relative reductions compared to the reference walls without weaknesses are essential for the subsequent MCS. For the simulation of the experiments, which are performed to validate the model, the compressive strength and the stress–strain parameter \( k \) are thus chosen as experimentally determined for the reference walls. The modulus of elasticity is set to the values obtained from the tests on RILEM specimens. Figure 7 illustrates the load–displacement curves from the experiments (average of the four LVDTs) and the corresponding simulations. The displayed loads are normalised regarding the cross-sectional area; that is, they are shown as average stress. The displacements are related to the length of the LVDTs and the height of the finite element wall, respectively. Therefore, they are displayed as an average strain. As evident, the

| Type of element                          | Parameter          | Value                  | References   |
|-----------------------------------------|--------------------|------------------------|--------------|
| Expanded units                          | General            |                        |              |
|                                         | Friction angle \( \varphi \) | 12°                    | Calibration  |
|                                         | Dilatancy angle \( \psi \)  | \( \varphi \)          | [22]         |
|                                         | Poisson’s ratio    | 0.19                   | [31]         |
|                                         | Compressive strength \( f_{ma} \) | 7.42 N/mm\(^2\) (single wythe) | Experiments |
|                                         |                    | 6.35 N/mm\(^2\) (cross bond) |              |
| Solid clay brick masonry (experiments)  | Modulus of elasticity \( E_{ma} \) | 3000 N/mm\(^2\)     | Experiments  |
|                                         | Stress–strain parameter \( k \) | 2.31 (single wythe) | Experiments  |
|                                         | Tensile strength \( f_{bt} \) | 0.04 \( f_b = 1.0 \text{ N/mm}^2 \) | [26]         |
|                                         | Compressive strength \( f_{ma} \) | 3.14 N/mm\(^2\) (single wythe) | Experiments  |
|                                         |                    | 2.64 N/mm\(^2\) (cross bond) |              |
| Perforated clay brick masonry (experiments) | Modulus of elasticity \( E_{ma} \) | 1959 N/mm\(^2\) | Experiments  |
|                                         | Stress–strain      | 1.69 (single wythe)    | Experiments  |
|                                         | Parameter \( k \)  | 2.04 (cross bond)      | [24] and [26]|
|                                         | Tensile strength \( f_{bt} \) | 0.03 \( f_b = 0.35 \text{ N/mm}^2 \) | [26]         |
| Solid clay brick masonry (Monte Carlo simulation) | Mean modulus of elasticity \( E_{ma} \) | 550 \( f_{ma}^{a} \) | [24] and [26]|
|                                         | Stress–strain      | 2.0                    | [24] and [25]|
|                                         | Parameter \( k \)  |                        |              |
|                                         | Tensile strength \( f_{bt} \) | 0.04 \( f_b = 0.11 \ f_{ma}^{a} \) | [26]         |
| Interfaces                              | (Mean) tensile strength \( f_t \) | 0.044 \( f_{ma}^{a} \) | [30]         |
|                                         | Tensile fracture energy \( G_{fn} \) | 0.0148 \( f_l \) | [31]         |

\(^a\)Relative material properties are based on \( f_b = 25 \text{ N/mm}^2 \), \( f_l = 2.5 \text{ N/mm}^2 \), and \( f_{ma} = 9.1 \text{ N/mm}^2 \)
simulated load–displacement curves match the experimental ones very well.

For the following MCS, only the differences between the experimental and simulated peak loads (i.e., the resistances) are important. Thus, the small differences in the post-peak behaviour are of minor relevance. The differences between the experimental and simulated results can—to some extent—be attributed to the inevitable variability of the experimental results themselves. To check the overall accuracy of the finite element model, the ratio $\frac{R_{\text{exp}}}{R_{\text{cal}}}$ between the average experimental resistance $R_{\text{exp}}$ and the simulated resistance $R_{\text{cal}}$ is calculated for each wall type with weaknesses. The average ratio $\theta = \frac{R_{\text{exp}}}{R_{\text{cal}}}$ is 1.03, which is very close to one. The CoV is 9.5%. Considering the variability of the experimental results themselves, this low CoV affirms that the numerical model is sufficiently precise for the MCS.

4 Stochastic modelling

4.1 General concept

The spatial variability of the material properties is modelled as unit-to-unit variability in the presented investigations. Masonry compressive strength $f_{\text{ma}}$, masonry elastic modulus $E_{\text{ma}}$, and joint tensile strength $f_t$ are represented by log-normal random variables. An individual random value of $f_{\text{ma}}$ and $E_{\text{ma}}$ is assigned to each expanded unit in the MCS. In the stochastic model, the interfaces at the bed joints are discretised into separate sections according to the units placed at their top, as the mortar is usually placed section by section for the next unit to be laid during construction. Each of these sections receives its individual tensile strength $f_{t,i}$. The compressive strength $f_{\text{ma},i}$ and the elastic modulus $E_{\text{ma},i}$ at one unit $i$ are correlated with the correlation coefficient $\rho_{f_{\text{ma}},E_{\text{ma}},i}$. The spatial correlation between the compressive strengths—as well as between the elastic moduli—at different expanded units $i$ and $j$ is defined by the correlation coefficient $\rho_{\text{spat}}$ (see Fig. 8).

According to this stochastic model, the correlation between the strengths of two expanded units in the wall is independent of their relative location to each other. The reason for the spatial correlation represented by $\rho_{\text{spat}}$ is that the units within one wall are likely to originate from the same production batch and are thus affected by common production conditions. Within the wall, however, the units from one construction batch are placed in an arbitrary order, which would make any location-related correlation between the unit strengths questionable. Although the compressive strength (and the elastic modulus of masonry) is influenced by the properties of both the units and the mortar, the described correlation structure is modelled according to that of the units. For justification, note that EN 1996-1-1 [32] defines the following empirical relationship between the characteristic masonry

![Stress–strain curves obtained from the experiments on masonry walls and corresponding finite element simulations](image-url)
compressive strength $f_{\text{ma},k}$ and the mean compressive strengths $f_{\text{b,m}}$ and $f_{\text{j,m}}$ of unit and mortar:

$$f_{\text{ma},k} = K f_{\text{b,m}}^{0.7} f_{\text{j,m}}^{0.3}$$  \hspace{1cm} (2)

Equation 2 corresponds to a linear relationship between the units of the logarithms of masonry, unit, and mortar compressive strengths. Hence, the resulting relationship between the respective variances $\sigma^2$ is

$$\sigma_{\ln f_{\text{ma}}}^2 = 0.72^2 \sigma_{\ln f_{\text{b}}}^2 + 0.32^2 \sigma_{\ln f_{\text{j}}}^2 = 0.49 \sigma_{\ln f_{\text{b}}}^2 + 0.09 \sigma_{\ln f_{\text{j}}}^2$$  \hspace{1cm} (3)

Equation 3 demonstrates the dominating influence of the unit properties on the variability of masonry compressive strength. This much stronger influence of the unit properties on masonry compressive strength justifies the choice of modelling correlation according to the unit properties. Analogous considerations apply to the modulus of elasticity.

4.2 Stochastic parameters

The CoVs of the material properties are varied in the parameter studies presented in Sect. 5. Therefore, typical ratios between the CoVs of the different material properties are specified rather than fixed CoVs. The CoVs of masonry compressive strength, masonry elastic modulus, and joint tensile strength are found to be 17%, 22%, and 35%, respectively, in experimental investigations by Schueremans [31]. Therefore, the relative CoVs of masonry elastic modulus and joint tensile strength are defined as 1.3 and 2 times the CoV of masonry compressive strength. These ratios also approximately match the CoVs specified in the JCSS Probabilistic Model Code [33]. The correlation between the compressive strength and the elastic modulus of masonry is set to $\rho_{f,E} = 0.72$, as evaluated in [31] based on experimental investigations. The spatial correlation coefficient $\rho_{\text{spat}}$ is varied between zero and one in the parameter studies.

4.3 Generation of random material properties

The generation of the random material properties masonry compressive strength and masonry elastic modulus according to the correlation structure in Fig. 8 and the stochastic parameters in Table 4 is performed according to an approach previously presented in [34]. The random values for $f_{\text{ma}}$ and $E_{\text{ma}}$ at a specific expanded unit $i$ are obtained as the product of four independent random variables:

$$f_{\text{ma},i} = W f_{\text{w},i} U_{f_{\text{w},i}}$$

$$E_{\text{ma},i} = W E_{\text{w},i} U_{E_{\text{w},i}}$$

where $W, E_{\text{w},i}, f_{\text{w},i},$ and $U_{i}$ are auxiliary random variables shared by particular pairs of the material properties $f_{\text{ma},i}$ and $E_{\text{ma},i}$. Since the same random values of $W$ and $U_{i}$ are applied to calculate $f_{\text{ma},i}$ and $E_{\text{ma},i}$, these random variables cause a correlation between the compressive strength and the modulus of elasticity at a specific unit $i$. In the same way, $W$ and $f_{\text{w},i}$ cause a correlation between the compressive strengths $f_{\text{ma},i}$ at two different units $i$ and $j$. Finally, the correlation between the moduli of elasticity $E_{\text{ma},i}$ at different units is caused by $W$ and $E_{\text{w},i}$. All random variables in Eqs. 4 and 5 are log-normally distributed. Furthermore, they have a mean value of one, except for $f_{\text{ma},i}$ and $E_{\text{ma},i}$, the mean values of which equal the actual mean values of $f_{\text{ma}}$ and $E_{\text{ma}}$. If the CoVs of the auxiliary variables are determined according to the following equations, the desired

| $f_{\text{ma},i}$ | $E_{\text{ma},i}$ | $f_{\text{ma},i}$ | $E_{\text{ma},i}$ | $f_{\text{ij}}$ | $f_{\text{ij}}$ |
|------------------|------------------|------------------|------------------|----------------|----------------|
| $f_{\text{ma},i}$ | 1                | $\rho_{f,E}$    | $\rho_{\text{spat}}$ | $\rho_{\text{spat}}$ | 0             | 0             |
| $E_{\text{ma},i}$ | 1                | $\rho_{\text{spat}}$ | $\rho_{f,E}$    | 0             | 0             |
| $f_{\text{ma},j}$ | 1                | $\rho_{f,E}$    | 0             | 0             |
| $E_{\text{ma},j}$ | 1                | 0             | 0             |
| $f_{\text{ij}}$ | 1                | 0             | 0             |
| $f_{\text{ij}}$ | $\text{symmetric}$ | 1             | 0             |

Fig. 8 Correlation between material properties
correlation coefficients $\rho_{f,E}$ and $\rho_{\text{spat}}$ and the CoVs $V_f$ and $V_E$ are obtained:

$$V_W = \sqrt{\rho_{\text{spat}} \rho_{f,E} V_E V_f}, \quad (6)$$

$$V_{f_w} = \sqrt{\frac{\rho_{\text{spat}} (V_f^2 - \rho_{f,E} V_E V_f)}{1 + \rho_{\text{spat}} \rho_{f,E} V_E V_f}}, \quad (7)$$

$$V_{U_i} = \sqrt{\frac{(1 - \rho_{\text{spat}}) \rho_{f,E} V_E V_f}{1 + \rho_{\text{spat}} \rho_{f,E} V_E V_f}}, \quad (8)$$

$$V_{f_{w,i}} = \sqrt{\frac{(1 - \rho_{\text{spat}}) (V_f^2 - \rho_{f,E} V_E V_f)}{(1 + \rho_{\text{spat}} V_E^2) (1 + \rho_{f,E} V_E V_f)}}, \quad (9)$$

The CoVs of $E_w$ and $E_{u,i}$ can be calculated analogously to those of $f_w$ and $f_{u,i}$ (i.e. by swapping the indices “$f$” and “$E$”). Equations 6 to 9 are derived from the well-known relationships for the variance of the product of random variables and the correlation coefficient of two random variables. Random material properties for the joint tensile strengths $f_{w,i}$ can be directly generated according to the desired mean and CoV, as they are not modelled as correlated.

### 5 Monte Carlo simulation

#### 5.1 Approach

The reference wall for the following parameter studies is illustrated in Fig. 9. It is arranged in cross bond with a thickness of two unit widths. Each of the 36 courses consists of five units of standard format NF [11], resulting in the height of 3 m, the thickness of 0.24 m, and the length of 0.625 m. The ratio of mean elastic modulus to mean compressive strength is set to $E_{\text{ma,m}} / f_{\text{ma,m}} = 550$ for the reference wall. The wall is loaded concentrically, and the supports are modelled as hinged around the longitudinal axis. The reference wall is simulated without consideration of geometrical nonlinearity (i.e. without second-order effects) since this leads to the most critical influence of spatial variability, as shown in the following. The CoV of masonry compressive strength is set to $V_f = 30\%$ for the reference wall, which is typical for existing solid clay brick masonry [2]. Based on the ratios in Table 4, the CoVs of masonry elastic modulus and joint tensile strength are 39% and 60%, respectively. Furthermore, no spatial correlation (i.e. full spatial variability) of the material properties and, hence, $\rho_{\text{spat}} = 0$ is assumed for the reference wall. In each parameter study in the following, one of the parameters of the reference wall is varied.

In the MCS, all material properties are normalised by the mean masonry compressive strength $f_{\text{ma,m}}$. In most parameter studies, 200 simulations with random material properties are conducted for each parameter combination. Subsequently, the mean $R_m$ and the CoV

| Material property       | Distribution type | Correlation | CoV in [31] (%) | Relative CoV |
|-------------------------|-------------------|-------------|-----------------|--------------|
| Masonry compressive strength $f_{\text{ma}}$ | Log-normal | $\rho_{f,E} = 0.72$ | 17              | 1            |
| Masonry modulus of elasticity $E_{\text{ma}}$ | Log-normal |             | 22              | $\approx 1.3$ |
| Tensile strength $f_t$ | Log-normal |             | 35              | $\approx 2.0$ |

![Fig. 9 Geometry and structural system of the reference wall](image-url)
$V_R$ of the resistance are estimated by the arithmetic mean and the sample CoV of the simulation results. Since lower quantile values are essential regarding structural reliability, design values of the resistance are also determined. As the focus is on existing structures, a corresponding target reliability index is chosen, which—due to higher costs of measures for increasing structural reliability—is typically lower than for the design of new structures [35, 36]. The target reliability index is selected as $\beta_{t,1a} = 3.3$ for a one-year reference period, as specified in ISO 2394 [37] for high relative costs of safety measures and medium failure consequences. Since the fixed sensitivity factor $\alpha_R = 0.8$ from EN 1990 is mainly defined for a 50-year reference period, a modified sensitivity factor $\alpha_{R,1a} = 0.7$ is used, which is more suitable for a one-year reference to achieve compatibility with design values based on a 50-year target reliability index $\beta_{t,50a}$ [38].

Although the input material properties are log-normally distributed, the resulting probability distribution of the wall resistance is not necessarily log-normally distributed. Anderson–Darling tests for goodness-of-fit [39] indicate that the most suitable distribution type varies between the parameter combinations. Nevertheless, all design values are evaluated assuming a log-normal distribution to gain better comparability. However, the respective variance is not estimated by the variance of the simulation results directly. To consider the skewness of the simulation results that might differ from that of a log-normal distribution, the distribution parameters $\mu_{\ln R}$ and $\sigma_{\ln R}^2$ (i.e. the mean and variance of the logarithm of the resistance) are chosen such that the mean and the 5% fractile of the distribution match the arithmetic mean and 5% fractile of the simulation results. The design value is then determined by

$$R_d = R_m \exp \left( -0.5 \sigma_{\ln R,5\%}^2 - \alpha_R \beta \sigma_{\ln R,5\%} \right)$$

$$= R_m \exp \left( -0.5 \sigma_{\ln R,5\%}^2 - 0.7 \cdot 3.3 \sigma_{\ln R,5\%} \right),$$

where $\sigma_{\ln R,5\%}^2$ denotes the variance derived from fitting the distribution parameters to the mean and 5% fractile of the results.

These design values are not directly suited for engineering practice, as model uncertainty has not been considered yet. The design values determined by Eq. 10 are compared to design values $R_{d,\text{hom}}$ obtained by the approach without considering spatial variability. Therefore, a deterministic calculation based on mean material properties is conducted first to obtain the deterministic resistance $R_{\text{det}}$ assuming homogeneity. According to the traditional approach without considering spatial variability, design values are determined based on the variability of the material properties. Here, masonry compressive strength $f_{ma}$ is the most relevant material property:

$$R_{d,\text{hom}} = R_{\text{det}} \exp \left( -0.5 \sigma_{\ln f}^2 - 0.7 \cdot 3.3 \sigma_{\ln f} \right)$$

with $\sigma_{\ln f}^2$ the variance of the logarithm of masonry compressive strength.

5.2 Influence of material variability

The evaluated simulation results for varying CoVs $V_{f}$ of masonry compressive strength are displayed in Fig. 10. As the spatial correlation coefficient $\rho_{\text{s,pat}}$ is zero, the CoV $V_f$ fully corresponds to unit-to-unit variability ($V_f = V_{f,\text{s,pat}}$). The resistance is normalised by cross-sectional area $A$ and mean masonry compressive strength $f_{ma,m}$. The normalised deterministic resistance $R_{\text{det}}/(A/f_{ma,m})$ is one since no load eccentricity and no second-order effects are considered.

Two main effects are evident. First, the mean resistance $R_m$ decreases with an increasing CoV of spatially variable masonry strength. This effect is
caused by an increasing number of expanded units
with very low compressive strength within the wall,
combined with a limited stress redistribution capabil-
ity. The mean resistance $R_m$ can be approximated as a
function of $V_{f,spat}$ as follows:

$$R_m \approx R_{m, appr} = R_{det} \exp(-1.62 V_{f,spat}^{1.44}).$$  (12)

Figure 10 demonstrates the excellent fit of the
approximation.

The second effect is that the resulting CoV of the
wall resistance $V_R$ is significantly smaller than the
input CoV of masonry compressive strength $V_f$. The
relationship between $V_f$ and $V_R$ can be approximated
by

$$V_R \approx V_{R, appr} = 0.17 V_{f,spat}$$  (13)

The CoV $V_{R,5%}$ corresponding to the variance
$\sigma_{ln R,5%}^2$ can analogously be approximated by 0.21
$V_{f,spat}$. All coefficients in the approximative equations
are determined using the method of least squares.

The reduction of the mean $R_m$ is a negative effect of
considering spatial variability, whereas the effect that
$V_R$ is much smaller than $V_f$ is positive. The overall
influence of considering spatial variability is apparent
when comparing the design value $R_d$ under consider-
ation of spatial variability to the design value $R_{d, hom}$
obtained by the traditional approach without spatial
variability. Figure 10 shows that the consideration of
spatial variability leads to significantly higher design
values here. Considering spatial variability thus has a
positive influence on obtained design values or partial
factors in this case.

5.3 Influence of spatial correlation

The spatial correlation coefficient $\rho_{spat}$ is varied
between 0 and 1 in this next parameter study. The
total CoV of masonry compressive strength is constant
with $V_f = 0.3$, but the choice of $\rho_{spat}$ determines which
part of the CoV corresponds to unit-to-unit and which
to wall-to-wall variability. The case $\rho_{spat} = 0$ is
identical to the reference case; all the variability of
masonry compressive strength belongs to spatial (i.e.
unit-to-unit) variability within the wall. For $\rho_{spat} = 1$,
the compressive strengths at the different units within
the wall are perfectly correlated. Hence, no spatial
variability exists within the wall; the CoV of masonry
compressive strengths fully corresponds to variability
between different walls. In the intermediate cases, the
variability consists of both a unit-to-unit and a wall-to-
wall component, represented by $V_{f,spat}$ and $V_{f,wall}$,
respectively. Based on Eqs. 6 to 9, these component
CoVs are given by

$$V_{f,wall} = V_f \sqrt{\rho_{spat}},$$  (14)

$$V_{f,spat} = V_f \sqrt{\frac{1 - \rho_{spat}}{1 + \rho_{spat} V_f^2}}.$$  (15)

In this parameter study, 100 simulations are
performed for each parameter combination. The
reduction of the simulation runs is possible, as the
random numbers for the auxiliary variables $W$, $E_w$, and
$f_w$, which are much more influential than those for the
individual units, are generated using Latin hypercube
sampling.

Figure 11 presents the MCS results for varying
values of $\rho_{spat}$. In the case of $\rho_{spat} = 1$ (i.e. homo-
genesis), the mean resistance equals the deterministic
resistance $R_{det}$. Furthermore, the CoV $V_R$ of the
resistance equals the CoV $V_f$ of masonry compressive
strength. Since this situation is identical to the
traditional approach for determining design values,
the design value $R_d$ derived from the MCS equals the
design value $R_{d, hom}$ (except for minor deviations due
to the limited number of simulation runs). As soon as
$\rho_{spat} < 1$, spatial variability is present, including its
previously described effects on the wall resistance,
namely the reduction of the mean $R_m$ and the CoV $V_{R}$. The positive effect of considering spatial variability on the design value is only present for a small correlation coefficient. For $\rho_{\text{spat}}$ around 0.5, the two effects of reduced mean and reduced variability approximately cancel each other out.

Based on Eqs. 11 and 12, the results for varying $\rho_{\text{spat}}$ can also be determined analytically. The reduction in $R_m$ is caused by only the spatial component of $V_f$, which can be determined with Eq. 14 to be inserted into Eq. 11. The CoV $V_R$ can be determined by considering that the resistance $R$ can be viewed as a product of two random variables: one representing a wall with $\rho_{\text{spat}} = 0$ and $V_f = V_{f,\text{spat}}$ and one for a wall with $\rho_{\text{spat}} = 1$ and $V_f = V_{f,\text{wall}}$. Since only the CoV of the first random variable is reduced according to Eq. 12, the CoV of the resistance is

$$V^2_R = 0.17^2 V^2_{f,\text{spat}} + V^2_{f,\text{wall}} + 0.17^2 V^2_{f,\text{spat}} V^2_{f,\text{wall}}.$$  

(16)

The CoV $V_{R,5\%}$ can be calculated analogously with the coefficient 0.21 instead of 0.17, enabling the determination of design values $R_d$. As illustrated in Fig. 11, the analytically determined values $R_{m,\text{appr}}$, $V_{R,\text{appr}}$, and $R_{d,\text{appr}}$ excellently match the simulation results. Hence, MCS with the finite element model are only required for $\rho_{\text{spat}} = 0$ since the intermediate cases can be captured analytically.

5.4 Influence of wall length

The number of units per course of the masonry wall is varied to study the influence of wall length. At first, a masonry pillar with just one undivided unit per course is investigated, followed by a pillar having two undivided units per course (with the next course always rotated by 90°). Then, walls in cross bond with three, five, seven, nine, and 11 units per course are investigated. The results are presented in Fig. 12. Although the number of potential weaknesses in the long walls is higher, the mean value $R_m$ increases with higher wall length. This is due to the higher number of units in a course involved in stress redistribution when the weakest unit in this course reaches its strength. Due to stress redistribution, the strengths of the expanded units are—to a certain extent—averaged out within a course, which leads to a lower CoV of the resistance for long walls. Because of these two positive effects, the design values significantly increase with higher wall length. Hence, this parameter study confirms the need for a reduction factor for the design value of walls with a small cross-sectional area $A$. In EN 1996-1-1 [32], such a reduction factor is specified as $0.7 + 3A$ for $A < 0.1\, \text{m}^2$, leading to 0.79, 0.88, and 0.97 for the investigated walls with one, two, and three units per course.

5.5 Influence of slenderness

This last parameter study aims at investigating the transition between material failure and stability failure. Therefore, geometrical nonlinearity is considered in this parameter study, and the wall slenderness is varied. In this context, it is helpful to define a material-related slenderness $\lambda$ that includes the ratio of elastic modulus $E_{\text{ma,m}}$ to compressive strength $f_{\text{ma,m}}$. Slenderness $\lambda$ is defined according to [40]:

$$\lambda = \frac{h_{\text{ef}}}{t} \sqrt{\frac{k f_{\text{ma,m}}}{E_{\text{ma,m}}}} = \frac{h_{\text{ef}}}{t} \sqrt{\frac{\varepsilon_{\text{f,ma}}}{k}},$$  

(17)

with $\varepsilon_{\text{f,ma}}$ the compressive strain at peak stress (based on the mean of $f_{\text{ma}}$ and $E_{\text{ma}}$), $k$ the stress–strain parameter, $h_{\text{ef}}$ the effective height of the wall length (i.e. buckling length), and $t$ the thickness of the wall. To avoid effects from changing the number or dimensions of the units, slenderness $\lambda$ is varied by altering the ratio $E_{\text{ma,m}}/f_{\text{ma,m}}$ between 10,000 and 75. With decreasing elastic modulus, second-order effects
gain more and more influence until the failure mode switches from compression to stability failure. The load eccentricity \( e \) is chosen as 0.1 \( t \) in this parameter study. Thereby, an initial eccentricity is present that is increased at the mid-height of the wall as a result of second-order effects. Due to hinged supports, \( h_{ef} \) equals the actual wall height \( h \). The results are presented in Fig. 13, where \( \lambda = 0 \) corresponds to a Monte Carlo simulation without considering geometrical nonlinearity. In addition to the mean \( R_m \), the CoV \( V_R \), and the design values \( R_d \) and \( R_{d,\text{hom}} \) of the resistance, the resistance \( R_{\text{det}} \) based on a deterministic simulation with mean material properties is shown, as \( R_{\text{det}}/(A \cdot f_{\text{ma},m}) \neq 1 \) here.

The influence of masonry compressive strength reduces with increasing slenderness \( \lambda \). In the case of high slenderness \( \lambda \), with most of the walls showing stability failure, only the overall stiffness, determined by the elastic moduli of the expanded units in the wall, is essential. While material failure is initiated by the failure of the weakest spots in the wall (i.e. the “weakest link”), the overall stiffness that determines the resistance in the case of stability failure is given by a weighted harmonic mean of the varying stiffness of the courses within the wall. Consequently, the gap between the mean resistance \( R_m \) and the deterministic resistance \( R_{\text{det}} \) gets smaller for increasing slenderness. Furthermore, the CoV \( V_R \) of the resistance is smaller in the case of high slenderness \( \lambda \), although the CoV of the elastic modulus is higher than the CoV of masonry compressive strength. As a result, the ratio \( R_d/R_{d,\text{hom}} \) and, hence, the positive effect of considering spatial variability is higher for slender walls showing stability failure than for non-slender walls with material failure.

6 Conclusions

The influence of spatially variable material properties on the resistance of solid clay brick masonry under compression has been investigated in the presented study. A finite element model was developed based on the simplified micro-modelling approach. The model proved its suitability by comparing experimental results on masonry walls featuring local weaknesses with corresponding finite element simulations. The finite element model was then utilised for Monte Carlo simulations, in which the spatial variability of masonry compressive strength, masonry elastic modulus, and joint tensile strength was modelled as unit-to-unit variability. Various parameter studies were performed to quantify the influence of spatial variability on the resistance of masonry under compression concerning the mean resistance, the CoV of the resistance, and the design value of the resistance. From these parameter studies, the following conclusions can be drawn:

- The mean resistance decreases with increasing spatial variability of the material properties within the wall. However, the CoV of the resistance is much smaller than the input CoV of the relevant spatially variable material property. Approximate equations were found for describing both effects.
- If the spatial correlation coefficient of the material properties within the wall is small, considering spatial variability usually leads to much higher design values of the resistance than the traditional approach that is based on the deterministic resistance assuming homogeneity and the variability of the most important material property. If the spatial correlation coefficient is around 0.5, both approaches approximately lead to the same design values for the investigated reference wall.
- Due to an improved stress redistribution capability, longer masonry walls perform better than short walls if spatial material variability is present.
- The relative reduction of the mean resistance caused by spatial variability is smaller for slender walls with stability failure, as average material
properties are more relevant than the “weakest link” in this case.

The results of the study thus provide valuable insights into the effects of spatial variability on the resistance of solid clay brick masonry under compression, which is particularly relevant for the assessment of existing masonry structures.

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Declarations

Conflict of interest The authors declare that they have no conflict of interest.

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