Which cuff should I use? Indirect blood pressure measurement for the diagnosis of hypertension in patients with obesity: a diagnostic accuracy review

Greg Irving, John Holden, Richard Stevens, Richard J McManus

ABSTRACT

Objective: To determine the diagnostic accuracy of different methods of blood pressure (BP) measurement compared with reference standards for the diagnosis of hypertension in patients with obesity with a large arm circumference.

Design: Systematic review with meta-analysis with hierarchical summary receiver operating characteristic models. Bland-Altman analyses where individual patient data were available. Methodological quality appraised using Quality Assessment of Diagnostic Accuracy Studies 2 (QUADAS2) criteria.

Data sources: MEDLINE, EMBASE, Cochrane, DARE, Medion and Trip databases were searched.

Eligibility criteria: Cross-sectional, randomised and cohort studies of diagnostic test accuracy that compared any non-invasive BP tests (upper arm, forearm, wrist, finger) with an appropriate reference standard (invasive BP, correctly fitting upper arm cuff, ambulatory BP monitoring) in primary care were included.

Results: 4037 potentially relevant papers were identified. 20 studies involving 26 different comparisons met the inclusion criteria. Individual patient data were available from 4 studies. No studies satisfied all QUADAS2 criteria. Compared with the reference test of invasive BP, a correctly fitting upper arm BP cuff had a sensitivity of 0.87 (0.79 to 0.93) and a specificity of 0.85 (0.64 to 0.95); insufficient evidence was available for other comparisons to invasive BP. Compared with the reference test of a correctly fitting upper arm cuff, BP measurement at the wrist had a sensitivity of 0.92 (0.64 to 0.99) and a specificity of 0.92 (0.85 to 0.87). Measurement with an incorrectly fitting standard cuff had a sensitivity of 0.73 (0.67 to 0.78) and a specificity of 0.76 (0.69 to 0.82). Measurement at the forearm had a sensitivity of 0.84 (0.71 to 0.92) and a specificity of 0.75 of (0.66 to 0.83). Bland-Altman analysis of individual patient data from 3 studies comparing wrist and upper arm BP showed a mean difference of 0.46 mm Hg for systolic BP measurement and 2.2 mm Hg for diastolic BP measurement.

Conclusions: BP measurement with a correctly fitting upper arm cuff is sufficiently sensitive and specific to diagnose hypertension in patients with obesity with a large upper arm circumference. If a correctly fitting upper arm cuff cannot be applied, an incorrectly fitting standard size cuff should not be used and BP measurement at the wrist should be considered.

INTRODUCTION

Approximately 671 million individuals worldwide are obese and subsequently mean arm circumference has increased. As a result, healthcare professionals are increasingly faced with situations where blood pressure (BP) measurement with a standard (or even large sized upper arm cuff) is not possible. This is important because errors in BP measurement are greater when the cuff used is too small relative to the patient’s arm circumference. In the USA, over 30% of the population have a large arm circumference as a consequence of obesity, a figure which can rise to over 60% in some clinics. Standard cuffs are ~22–32 cm with large cuffs typically ranging from 32 to 42 cm.

Strengths and limitations of this study

- Study quality was assessed using Quality Assessment of Diagnostic Accuracy Studies 2 (QUADAS2).
- Forest plots of sensitivity and specificity were created for each study and estimates of sensitivity and specificity in receiver operating characteristic space for each comparison were plotted.
- Individual patient data were obtained and the Bland-Altman method used to plot differences in measurement against reference standard.
- Insufficient evidence was available to compare forearm, wrist, finger blood pressure measurement with invasive blood pressure.
- All included studies used body mass index (BMI) >30 as an indicator of obesity. Although BMI generally increases as adiposity increases, due to variation in body composition, it is not the most accurate measure of body adiposity.
Where available cuffs are not large enough for an individual’s upper arm, clinicians may use a variety of different methods for BP measurement including a correctly fitting extra large cuff, forearm BP, wrist BP, finger BP or ambulatory BP. However, there is considerable uncertainty as to the diagnostic accuracy of these different approaches for obese people and hence the optimum alternative test. Currently, all of the major hypertension guidelines recommend different approaches (table 1).

Understanding which alternative method to use is important because hypertension is common in patients with obesity and the consequences of missing the diagnosis is potentially severe at the individual and population level. This review aimed to determine the diagnostic accuracy of upper arm, forearm, wrist and finger BP measurement compared with the reference standards of invasive arterial, upper arm (correctly fitting cuff) and/or ambulatory BP measurement (ABPM) for the diagnosis of hypertension in patients with obesity with a large arm circumference.

### Table 1 Implications for policy

| Guideline | Guidance | Comment |
|-----------|----------|---------|
| American Heart Association | “In patients with morbid obesity, one will encounter very large arm circumferences…In this circumstance, the clinician may measure blood pressure from a cuff placed on the forearm and listening for sounds over the radial artery or use a validated wrist blood pressure monitor held at the level of the heart.” | There was insufficient evidence to undertake subgroup analyses for patients with an arm circumference >50 cm. However, on the basis of the available evidence in this review, if a correctly fitting upper arm cuff cannot be placed then measurement at the wrist is preferable to measurement at the forearm. |
| European Society of Cardiology | “Devices worn on the wrist are currently not recommended but their use might be justified in obese subjects with extremely large arm circumference.” | This review would support the use of wrist BP devises in patients with obesity if a correctly fitting upper arm cuff cannot be applied and the device was used at the level of the heart. |
| British Hypertension Society/ National Institute for Clinical Excellence | “A large cuff is still too small. What should I do? Contact the manufacturer of the blood pressure monitor. They may be able to supply an extra large cuff.” No specific guidance is offered for the use of wrist BP measurement in patients with obesity with a large arm circumference. “The bladder of the cuff should fit around at least 80% of the arm but not more than 100%. A cuff that does not fit properly will not give an accurate reading so it is important to use the right size.” | Compared with the reference test of invasive BP, a correctly fitting upper arm BP cuff is sufficiently sensitive and specific for the diagnosis of hypertension in patients with obesity. This holds true for patients with BMI>35. It was not possible to undertake a subgroup analysis of those with arm circumference >40 or 50 cm. There was good evidence that an incorrectly fitting standard cuff was not as accurate as a correctly fitting cuff. Both wrist and forearm BP measurements were more accurate than an incorrectly fitting upper arm cuff. |
| WHO/International Society for Hypertension | No specific guidance is offered for the use of wrist BP measurement in patients with obesity with a large arm circumference. | Compared with the reference test of invasive BP, a correctly fitting upper arm BP cuff is sufficiently sensitive and specific for the diagnosis of hypertension in patients with obesity. This holds true for patients with BMI>35. It was not possible to undertake a subgroup analysis of those with arm circumference >40 or 50 cm. |

BMI, body mass index; BP, blood pressure.
and both manual and automated BP measurements were considered. Obese adults were defined by either an upper arm circumference ≥35 cm, a body mass index (BMI) ≥30 or by direct measurement of percentage body fat (≥25% for men and ≥30% for women). Prespecified subgroups included those with arm circumference ≥40 cm or ≥50 cm or BMI ≥35.

Included study designs were diagnostic cross-sectional, randomised and cohort studies. Studies were excluded if participants were: receiving antihypertensive treatment at the time of comparison or pregnant or a hospital inpatient. Studies from which data could not be extracted were included in the descriptive part of the review but excluded from subsequent analyses.

**Selection and data extraction**

To determine inclusion or exclusion of each potential study, GI and JH independently reviewed the results of the search by title/abstract and when required by the full text of the study. The resulting list of citations was then reviewed for inclusion and a third author (RJM) arbitrated the final selection decisions.

A standardised data extraction form was used to identify study characteristics and results for each included publication which were independently extracted by GI and JH. Study quality was assessed using Quality Assessment of Diagnostic Accuracy Studies 2 (QUADAS2). Where possible, data were extracted into 2×2 contingency tables. All authors were contacted directly if insufficient data were available in the published report of a given study and to request individual patient data.

**Statistical analysis and data synthesis**

Forest plots of sensitivity and specificity were created for each study and estimates of sensitivity and specificity in receiver operating characteristic (ROC) space for each comparison were plotted. The diagnostic performance of each index test was ascertained and heterogeneity in such performance estimated (see below). The Metandi and Midas procedures in STATA V13.0 were used to fit the hierarchical summary receiver operating curve (HSROC) models (R Harbord. METANDI: Stata module to perform meta-analysis of diagnostic accuracy. Statistical Software Components 2008; Stata Corporation. Stata Statistical Software Release V13.0: Programming; Stata Corporation, 2001). Where there was a sufficient number of studies (n≥5) the HSROC model was used to derive inferences about diagnostic test accuracy and heterogeneity in test performance including the summary curve. Where individual patient data were available, the Bland-Altman method was used to plot differences in measurement against reference standard.

**Investigations of heterogeneity and sensitivity analyses**

Heterogeneity was evaluated by visual inspection of Forest plots of sensitivity and specificity, ROC plots of the data and using the chi-squared test. The influence of differences in the test characteristics (manual vs automated BP device), study population (primary care vs hospital outpatient) and methodological quality (scores on items of the QUADAS2 checklist) was investigated where it was appropriate to do so.

Where there was evidence of differences between studies sensitivity analyses were undertaken based on:

1. Measurement of obesity: indirect/direct methods, for example, bioelectrical impedance analysis (BIA) or skin fold thickness.
2. Risk of bias according to QUADAS2 criteria: low/high risk of bias.
3. Year of study: <25/≥25 years old.

Reporting bias was assessed using the test for funnel plot asymmetry.

**Patient involvement**

The design of the review was informed by a search of patient uncertainties in the UK Database of Uncertainties of the Effects of Treatment (DUETS) and patient research priorities identified by the James Lind Alliance.

**RESULTS**

**Results of the search**

A total of 4037 studies were identified (excluding duplicates). The full text of 164 papers was reviewed for eligibility (figure 1) and 37 studies were included (32 published in English). One was a randomised trial and one a case–control study; the remainder were cross-sectional studies. Twenty studies (26 comparisons) had extractable data, all of which were of cross-sectional design. There were no disagreements between authors in relation to the number of studies eligible for inclusion (κ=1.0). Individual patient data were available from four authors.

**Methodological quality of included studies**

Quality assessment of the 20 included studies and the 16 studies that could not contribute data, found no studies that satisfied all the QUADAS2 criteria as all had some degree of methodological weakness and/or lacked reporting clarity (figure 2). Blinding of reference tests, blinding of index tests and acceptable delay between tests were particularly poorly reported.

**Findings**

Six studies were found which used invasively measured BP as the reference standard and a further 17 using a properly fitting upper arm cuff. One study used ABPM as a reference standard. Details of included studies are shown in table 2 and results are presented by reference test below.

**Reference test: invasively measured BP**

Six studies were found with extractable data (eight data sets) comparing upper arm BP with invasive
measurement in obesity yielding a pooled sensitivity of 0.87 (0.79 to 0.93) and a specificity of 0.85 (0.64 to 0.95). Only three studies were inside the 95% CIs for the summary receiver operating curve (SROC) summary point. The five outlying studies had a relatively small sample size (5–20 participants). Deeks funnel plot asymmetry test had a non-significant p value (p=0.53) for the slope coefficient suggesting no evidence of publication bias. Only one study was included for the subgroup analysis of those patients with BMI>35. This resulted in a sensitivity of 0.88 (0.47 to 1.00) and a specificity of 0.71 (0.40 to 0.90). There were insufficient data to perform any other subgroup analyses. Across the six studies there was no difference in reference threshold and all avoided verification bias.

**Figure 1** Date of search 1/6/2015. PRISMA flow diagram. DTA, diagnostic test accuracy.

**Figure 2** Risk of bias and applicability concerns summary with review authors’ judgements for each included study.
There was no significant evidence of heterogeneity across the six studies ($\chi^2: Q=1.51, df=2, p=0.24$). Exclusion of the two studies that were >25 years old (1953 and 1965) and used older equipment resulted in a pooled sensitivity of 0.86 (0.77 to 0.92) and a specificity of 0.90 (0.70 to 0.92) and did not influence heterogeneity. The number of studies was too small to calculate a pooled analysis for other a priori stated potential sources of heterogeneity.

No studies with extractable data regarding forearm, wrist or finger BP in obese individuals compared with invasively measured BP were found.

**Reference test: correctly fitting upper arm cuff**

Sixteen studies (18 comparisons) were found with a reference standard of correctly fitting upper arm cuff compared with a standard cuff, forearm, wrist or finger BP measurement.

**Upper arm BP**

Five studies (six data sets) compared a standard sized (ie, too small) upper arm cuff with a correctly fitting upper arm cuff and found a pooled sensitivity of 0.73 (0.67 to 0.78) and a specificity of 0.76 (0.69 to 0.82). The majority of the studies were within the 95% CI of the SROC summary point. The two outliers had a relatively small sample size in comparison to the other studies (n=4 and n=31) and the Deeks funnel plot suggested no evidence of publication bias (p=0.34). Only one study was included in the subgroup analysis of those patients with BMI>35 with resultant wide CIs: sensitivity 1.00 (0.59 to 1.00) and

---

**Table 2 Characteristics of included studies**

| Study          | Sample size (patients with obesity) | Mean age | Male (%) | Study population | Study design  |
|----------------|-------------------------------------|----------|----------|------------------|--------------|
| Bennett et al  | 26                                  | –        | –        | Community        | Cross-sectional |
| Bernsten et al | 4                                   | 65       | –        | Community        | Cross-sectional |
| Bertrand et al | 16                                  | –        | –        | Community        | Case-control   |
| Blackburn 1965 | 76                                  | –        | 100      | Hospital outpatient | Cross-sectional |
| Bovet 1994     | 103                                 | –        | 43.7     | Hospital outpatient | Cross-sectional |
| Chiolero et al | 62                                  | –        | 49.7     | Community        | Cross-sectional |
| Cuckson et al  | 20                                  | –        | –        | Hospital outpatient | Cross-sectional |
| de Senarclens et al | 12                              | 39.9 | 27      | Hospital outpatient | Cross-sectional |
| Domiano et al  | 42                                  | 50.7     | 56       | Community        | Cross-sectional |
| Doshi et al    | 103                                 | 39.9     | 27       | Hospital outpatient | Cross-sectional |
| Genc et al     | 32                                  | 51.4     | 44.8     | Hospital outpatient | Cross-sectional |
| Guagnano et al | 97                                  | 48.7     | 35       | Hospital outpatient | Cross-sectional |
| Guagnano et al | 94                                  | 42.9     | 0        | Hospital outpatient | Cross-sectional |
| Guagnano et al | 339                                 | 41       | 0        | Hospital outpatient | Cross-sectional |
| Julien 1983    | 19                                  | 51       | 5        | Hospital outpatient | Cross-sectional |
| King 1967      | –                                   | –        | –        | Hospital outpatient | Cross-sectional |
| Kotis 2005     | 825                                 | 53.5     | 53       | Hospital outpatient | Cross-sectional |
| Kvals 1969     | –                                   | –        | –        | Hospital outpatient | Cross-sectional |
| Leblanc et al  | 25                                  | –        | –        | Hospital outpatient | Cross-sectional |
| Linfors et al  | 116                                 | 51       | 29       | Hospital outpatient | Cross-sectional |
| Maxwell 1982   | –                                   | –        | –        | Hospital outpatient | Cross-sectional |
| Maxwell 1985   | –                                   | –        | 51       | Hospital outpatient | Cross-sectional |
| Nielsen et al  | 57                                  | 41       | 53       | Hospital outpatient | Cross-sectional |
| Palatini et al | –                                   | 54       | 50       | Hospital outpatient | Cross-sectional |
| Pang 2006      | –                                   | –        | –        | Hospital outpatient | Cross-sectional |
| Pielin et al   | 155                                 | 45       | 10       | Hospital outpatient | Cross-sectional |
| Poncelet et al | 6                                   | 52       | 33       | Hospital outpatient | Cross-sectional |
| Schell 2005    | 204                                 | 36.5     | 51.5     | Hospital outpatient | Cross-sectional |
| Schell 2006    | 79                                  | –        | 46.6     | Hospital outpatient | Cross-sectional |
| Simpson et al  | 5                                   | –        | –        | Hospital outpatient | Cross-sectional |
| Stergiou et al | 20                                  | –        | –        | Community        | Randomised trial |
| Stolt et al    | 10                                  | –        | –        | Community        | Cross-sectional |
| Stolt et al    | 20                                  | –        | 48       | Community        | Cross-sectional |
| Umana et al    | 61                                  | 62       | 62       | Hospital outpatient | Cross-sectional |
| van Montfrans 1987 | 19                              | –        | –        | Hospital outpatient | Cross-sectional |
| Vinyoles et al | 108                                 | 60.2     | 22.2     | Community        | Cross-sectional |
| Warembourg et al | 10                              | 54       | 0        | Hospital outpatient | Cross-sectional |

*Studies with extractable data.

All studies used body mass index ≥30 as definition of obesity.
specificity 0.67 (0.22 to 0.96). There were insufficient data to perform any other subgroup analyses.

Across the five studies, there was no difference in reference threshold and all studies avoided verification bias. There was no significant statistical evidence of heterogeneity across the six comparisons (\(\chi^2: Q=1.62, df=2, p=0.22\)) but there were insufficient data to carry out a pooled analysis for other potential sources of heterogeneity.

**Forearm BP**

Six studies compared BP measurement at the forearm at the level of the heart with a correctly fitting upper arm cuff and found a sensitivity of 0.84 (0.71 to 0.92) and a specificity of 0.75 (0.66 to 0.83). Four of the studies were within the 95% CI of the SROC summary point. Deeks funnel plot asymmetry test had a non-significant p value (p=0.50) for the slope coefficient suggesting no evidence of publication bias. There was statistical evidence of heterogeneity across the five studies (\(\chi^2: Q=8.382, df=2.00, p=0.008\)) which appeared largely due to two studies with much earlier publication dates. Excluding these studies resulted in a pooled sensitivity of 0.86 (0.77 to 0.92) and a specificity of 0.90 (0.70 to 0.92) and eliminated heterogeneity (\(\chi^2: Q=0.68, df=2.00, p=0.36\)). Further analyses for heterogeneity or subgroups were not possible.

**Wrist BP**

Five studies considered BP measured at the wrist held at the level of the heart compared with a properly fitting upper arm cuff yielding a pooled sensitivity of 0.92 (0.64 to 0.99) and a specificity of 0.92 (0.85 to 0.97). Funnel plot and \(\chi^2\) tests suggested a low probability of publication bias (p=0.89) or heterogeneity (\(\chi^2: Q=4.26, df=2.00, p=0.06\)), and subgroup analyses were not possible.

**Finger BP**

One study with two extractable data sets compared finger BP measurement with that from a correctly fitted upper arm cuff and reported a similar specificity (0.57 and 0.61) but a markedly different sensitivity (0.74 and 0.91) from the two comparisons. With so few data, it was not possible to calculate a pooled summary estimate, assess for heterogeneity or consider subgroups or publication bias.

**Comparing methods**

A summary ROC plot of all index tests compared with the reference test of a correctly fitting upper arm cuff is provided in figure 3. Visual inspection suggests that measurement at the wrist held at the level of the heart performs best in comparison to the reference.
Individual patient data
Bland-Altman analysis of the available individual patient data for three studies comparing wrist and upper arm BP measurement with a correctly fitting cuff showed a mean difference of 0.51 mm Hg (limits of agreement −21.7 to 22.6 mm Hg; systolic BP) and 1.96 mm Hg (limits of agreement −14.1 to 18.6 mm Hg; diastolic BP; table 3).17 21 23

Reference test: ambulatory BP
One study with extractable data compared upper arm BP measurement with that of ambulatory BP in obesity.25 This showed a sensitivity of 0.45 (0.29 to 0.62) and a specificity of 0.71 (0.60 to 0.81). Given that only a single study was identified, we were not able to assess for statistical heterogeneity or undertake subgroup analyses. No studies with extractable data regarding forearm, wrist or finger BP in obese individuals compared with ambulatory monitoring were found.

DISCUSSION
Summary of main results
This review has shown that measurement of BP in an obese individual using a correctly fitting upper arm BP cuff compared with the reference standard of direct arterial measurement provides similar results to those obtained in non-obese adults: sensitivity (0.87 (0.79 to 0.93; obese)) versus (0.67 (0.30 to 0.93; non-obese)) and specificity (0.85 (0.64 to 0.95; obese) versus (0.95 (0.83 to 0.99; non-obese); table 4).34 There is insufficient evidence to comment on how forearm, wrist and finger BP testing perform in relation to invasively measured BP in obesity.

Although CIs around the point estimates overlapped, inspection of SROC curves suggested BP measurement at the wrist is the alternative of choice should a correctly fitting upper arm cuff not be available. Bland-Altman analysis of available individual patient data demonstrated good agreement between wrist and upper BP measurement with a correctly fitting cuff for systolic (mean difference 0.51 mm Hg) and diastolic BP (mean difference 1.96 mm Hg). This falls within the ±3 mm Hg (British Hypertension Society (BHS) standard) and ±5 mm Hg (clinically relevant difference) margin of error. However, it is important to note that this analysis was based on a relative small amount of data with wide limits of agreement and that care was taken to minimise the influence of arm–heart hydrostatic pressures by asking the patient to hold the wrist devices at the level of the heart. If this is not performed, then systematic errors in measurement are likely to occur.6

Limitations of the review
For a number of the studies included in this review, data were either missing or could not be extracted directly from the published paper in order to construct the

Table 3  Bland-Altman mean difference values for individual studies of wrist versus correct cuff

| Study              | Specificity | Systolic BP mean difference (limits of agreement±2 SD) | Diastolic BP mean difference (limits of agreement±2 SD) |
|--------------------|-------------|-----------------------------------------------------|------------------------------------------------------|
| Chiolero et al21    | 0.93        | 0.51 (−22.7 to 23.7)                                 | 1.96 (−16.3 to 20.2)                                 |
| Doshi et al26       | 0.91        | 2.61 (−20.9 to 15.7)                                 | 4.98 (−18.1 to 8.2)                                  |
| de Senarcens et al23| 0.91        | −6.09 (−22.3 to 10.2)                                | 1.36 (−12.6 to 15.3)                                 |
| Stergiou et al27    | 0.88        | 4.05 (−14.8 to 22.9)                                 | 3.61 (−6.0 to 13.2)                                  |
| All*               |             | 0.46 (−22.7 to 22.6)                                 | 2.2 (−14.1 to 18.6)                                  |

* Doshi et al26 not included in pooled result as no individual patient data available. The figure presented is taken from the Bland-Altman analysis provided in the paper.

BP, blood pressure.

Table 4  Summary of results

| Reference Index          | Number of Studies (number of participants) | Sensitivity (95% CI) | Specificity (95% CI) | Positive likelihood ratio (95% CI) | Negative likelihood ratio (95% CI) |
|--------------------------|-------------------------------------------|----------------------|----------------------|------------------------------------|-----------------------------------|
| Invasive                 | Correctly fitting cuff                     | 6 (163)              | 0.87 (0.79 to 0.93)  | 0.85 (0.64 to 0.95)                | 3.33 (1.89 to 5.88)               | 0.18 (0.09 to 0.35)               |
| Correctly fitting cuff   | Standard cuff                              | 6 (621)              | 0.73 (0.67 to 0.78)  | 0.76 (0.69 to 0.82)                | 2.74 (2.03 to 3.70)               | 0.38 (0.30 to 0.47)               |
| Correctly fitting cuff   | Forearm cuff                               | 6 (425)              | 0.84 (0.71 to 0.92)  | 0.75 (0.66 to 0.83)                | 3.64 (2.47 to 5.37)               | 0.17 (0.08 to 0.34)               |
| Correctly fitting cuff   | Wrist cuff                                 | 5 (217)              | 0.92 (0.64 to 0.99)  | 0.92 (0.85 to 0.87)                | 12.62 (5.90 to 27.0)              | 0.82 (0.01 to 0.49)               |
required contingency tables. Authors were contacted directly to obtain these data, and individual patient data made available by three researchers was used to calculate Bland-Altman analyses.

All included studies used BMI>30 as an indicator of obesity which has limitations. Although BMI generally increases as adiposity increases, due to variation in body composition, it is not the most accurate measure of body adiposity. Those with a high proportion of muscle mass could have a high BMI. In a similar way arm circumference as a measure of arm obesity may be confounded in the case of an athlete with a low percentage body fat and large, muscular arms. No studies used direct methods such as BIA, skin fold thickness, underwater weighing. In some populations, the percentage of patients without obesity with a large arm circumference arms can be ~10% (Latman 2013, personal communication).

Some statistical heterogeneity was apparent (forearm vs correctly fitting upper arm cuff) and this could not be explained by reference threshold differences or partial verification. Instead, other factors are likely to have contributed, in particular publication date, suggesting the quality of equipment used may be important. However, it is possible other sources may have contributed including population characteristics (prevalence of hypertension) or test application (operator variability). Where heterogeneity was not detected, it is important to be aware of the limitations of tests for heterogeneity in diagnostic accuracy studies particularly when the number of included studies is small.

Comparison with the literature

This review is the first, to the best of our knowledge, to summarise the international literature on indirect BP measurement in obesity. This is perhaps reflected in the variation between international guidelines in their recommendations for BP measurement in obesity. The American Heart Association currently advocate the use of forearm BP; the European Society of Cardiology advocate wrist BP measurement and the British Hypertension Society suggest contacting manufacturers in order to obtain an extra large upper arm cuff. The International Society of Hypertension offers no specific advice.

Clinical and policy implications

On the basis of the review, compared with the reference test of invasive BP, a correctly fitting upper arm BP cuff is sufficiently sensitive and specific for the diagnosis of hypertension in patients with obesity. This also hold true for patients with a BMI>35. If a correctly fitting cuff cannot be fitted or is unavailable, then wrist BP measurement appears to be the next best alternative. BP measurement with an incorrectly fitting standard size cuff on the upper arm should be avoided as it is an insufficiently sensitive or specific test for a diagnosis of hypertension.

All included studies had some degree of methodological weakness and/or lacked of clarity in their reporting. Blinding of reference tests, blinding of index tests and acceptable delay between tests were particularly poorly reported. For this reasons, it is important for any future studies to address these important methodological considerations and follow the QUADAS2 guidance. On the basis of the paucity of individual patient data, there is now a need for a large diagnostic accuracy study of high methodological quality comparing wrist/forearm BP measurement with upper arm BP. This should aim to include patients with an arm circumference at least 40 cm. To determine percentage arm adiposity more accurately direct methods such as BIA and skin fold thickness could be employed.

CONCLUSIONS

This review set out to identify the diagnostic accuracy of non-invasive BP measurement techniques (upper arm, forearm wrist, finger) for the diagnosis of hypertension in patients with obesity with a large arm circumference compared with three different reference standards: invasive BP, correctly fitting upper arm cuff and ABPM. In conclusion, in the absence of a correctly fitted upper arm cuff, measurement at the wrist appears appropriate, provided that the cuff is held at the level of the heart.

Author affiliations

1Department of Primary Care, Institute of Public Health and Primary Care, University of Cambridge, Cambridge, UK
2Department of Garswood Surgery, Garswood, UK
3Nuffield Department of Primary Care Health Sciences, NIHR School for Primary Care Research, University of Oxford, Oxford, UK

Acknowledgements The authors would like to thank those authors who shared study data (Bennett, Berntsen, Bertrand, Chioro, Cuckson, de Senarclens, Domiano, Doshi, Guagnano, Leblanc, Linfors, Nielsen, Pierin, Poncelet, Schell, Stergioul, Stoit, Stolt, Vinyoles).

Contributors GI and RJM conceived and designed the review, GI and JH screened the references and assessed risk of bias. GI, RS and RJM analysed the data. All authors revised and approved the final version of the review.

Funding GI is funded as an NIHR Clinical Lecturer in General Practice.

Competing interests None declared.

Provenance and peer review Not commissioned; externally peer reviewed.

Data sharing statement No additional data are available.

Open Access This is an Open Access article distributed in accordance with the Creative Commons Attribution Non Commercial (CC BY-NC 4.0) license, which permits others to distribute, remix, adapt, build upon this work non-commercially, and license their derivative works on different terms, provided the original work is properly cited and the use is non-commercial. See: http://creativecommons.org/licenses/by-nc/4.0/

REFERENCES

1. Ng M, Fleming T, Robinson M, et al. Global, regional, and national prevalence of overweight and obesity in children and adults during 1980–2013: a systematic analysis for the Global Burden of Disease Study 2013. Lancet 2014;384:766–81.
2. Palatini P, Parati G. Blood pressure measurement in very obese patients: a challenging problem. J Hypertens 2011;29:425–9.
3. Doshi H, Weder AB, Bard RL, et al. Does “hidden undercuffing” occur among obese patients? Effect of arm sizes and other
