Joint large deviation result for empirical measures of the coloured random geometric graphs

Kwabena Doku-Amponsah

Abstract
We prove joint large deviation principle for the empirical pair measure and empirical locality measure of the near intermediate coloured random geometric graph models on \( n \) points picked uniformly in a \( d \)-dimensional torus of a unit circumference. From this result we obtain large deviation principles for the number of edges per vertex, the degree distribution and the proportion of isolated vertices for the near intermediate random geometric graph models.

Keywords: Random geometric graph, Erdős–Rényi graph, Coloured random geometric graph, Typed graph, Joint large deviation principle, Empirical pair measure, Empirical measure, Degree distribution, Entropy, Relative entropy, Isolated vertices

Mathematics Subject Classification: 60F10, 05C80, 68P30

Background
In this article we study the coloured geometric random graph CGRG, where \( n \) points or vertices or nodes are picked uniformly at random in \([0, 1]^d\), colours or spins are assigned independently from a finite alphabet \( \Sigma \) and any two points with colours \( a_1, a_2 \in \Sigma \) distance at most \( r_n(a_1, a_2) \) apart are connected. This random graph models, which has the geometric random graph (see Penrose 2003) as special case, has been suggested by Cannings and Penman (2003) as a possible extension to the coloured random graph studied in Biggins and Penman (2009), Doku-Amponsah and Mörters (2010), Doku-Amponsah (2006, 2014), Bordenave and Caputo (2015), Mukherjee (2014) and Doku-Amponsah (2014).

The connectivity radius \( r_n \) plays similar role as the connection probability \( p_n \) in the coloured random graph model. Several large deviation results about the coloured random graphs and hence Erdős–Rényi graph have been established recently. See O’Connell (1998), Biggins and Penman (2009), Doku-Amponsah and Mörters (2010), Doku-Amponsah (2006, 2014), Bordenave and Caputo (2015), Mukherjee (2014).

Until recently few or no large deviation result about the CGRG have been found. Doku-Amponsah (2015) proved joint large deviation principle for empirical pair measure and the empirical locality measure of the CGRG, where \( n \) points are uniformly chosen in \([0, 1]^d\), colours or spins are assigned by drawing without replacement from the
pool of, say, \( n_{\nu}(a_1) \) colours, and \( n_{\omega}(a_1, a_2) \) edges, \( a_1, a_2 \in \Sigma \), are randomly inserted among the points for some colour law \( \nu_n: \Sigma \to [0, 1] \) and edge law \( \omega_n: \Sigma \times \Sigma \to [0, \infty) \).

This article presents a full joint large deviation principle (LDP) for the empirical pair measure and the empirical locality measure of the CGRG. Refer to (Doku-Amponsah and Mörters) for similar result for the coloured random graphs. From this large deviation results we obtain LDPs for graph quantities such as number of edges per vertex, the degree distribution and the proportion of isolated vertices of geometric random graphs in the intermediate case. Our results are similar to those in O’Connell (1998), Biggins and Penman (2009), Doku-Amponsah and Mörters (2010), Doku-Amponsah (2006, 2014), Bordenave and Caputo (2015) and Mukherjee (2014) for the Erdő–Rényi graph except that the rate functions of the LDPs in our current setting is bigger as a result of the effect of the geometric in the model.

As a first step in the proof of our main result, we obtain a joint LDP for the empirical colour measure and empirical pair measures for the CGRG, see Theorem 4, by the exponential change-of-measure techniques and coupling argument. See example Doku-Amponsah and Mörters (2010) or Doku-Amponsah (2006). In the next step, we use Biggins (2004, Theorem 5(b)) to mix Theorem 4 and the result (Doku-Amponsah 2015, Theorem 2.1) to obtain the full joint LDP for empirical pair measure and the empirical locality measure of CGRG model. Refer to Doku-Amponsah and Mörters (2010) or Doku-Amponsah (2006) for further illustration of this method.

Our main motivation for studying this model are in two folds.

**Independence testing**  Consider CGRG which is a model for Wireless Sensor Network as a very big dataset comprising the typed sites and the bonds between sites. One interesting question to ask is how many bits will be required to code the \( n \) sites and the bonds between sites with high probability? Then, an asymptotic equipartition property (AEP) for the WSN will answer this question and our LDP for the empirical measures of the CGRG will play a crucial in the prove of the AEP. Further, we can test whether a received codeword \( y_n \) of WSN is jointly typical with a candidate sent codeword \( x_n \) of WSN. The probability that two independent sequences \((x_n, y_n)\) (\( x_n \) being a codeword other than what was sent when \( y_n \) was received) actually appear as dependent is bounded asymptotically as \( 2^{-n I} \), where the AEP is used to obtain the bound. See Doku-Amponsah (2016) for more on this application.

**Hypothesis testing**  One of the standard problems in statistics is to decide between two alternative explanations for the data are observed. For example, a transmitter will send an information on the WSN bits by bits in communication systems. There are two possible cases for each transmission: one is that bit 0 of WSN data is sent (noted as event \( H_0 \)) and the other is that bit 1 of WSN data is sent (noted as event \( H_1 \)). In the receiver side, the bit \( y \) is to be received as either 0 or 1. Based on the \( y \) bit of WSN data received, we can make a hypothesis whether the event \( H_0 \) happens (bit 0 was sent at the transmitter) or the event \( H_1 \) happens (i.e. bit 1 was sent at the transmitter). Of course, we may make mis-judgement, such as we decode that bit 0 was sent but actually bit 1 was sent. We need to make the probability of error in hypothesis testing as low as possible and the LDPs for CGRG models can help us specify the probability of error.
In the remainder of the paper we state and prove our LDP results. In “Statement of the results” section we state our LDPs, Theorem 1, Corollary 2, Corollary 3, Theorem 4, and Corollary 5. In “Proof of Theorem 1” section we present the proof of Theorem 4. In “Proof of Theorem 1” section we combine Theorem 2.1 and Doku-Amponsah (2014, Theorem 2.1) to obtain the Theorem 1, using the setup and result of Biggins (2004) to ‘mix’ the LDPs. The paper concludes with the proofs of Corollaries 2, 3 and 5 which are given in “Proof of Corollaries 2, 3, and 5” section.

**Statement of the results**

The joint LDP for empirical pair measure and empirical locality measure of CGRG

In this subsection we shall look at a more general model of random geometric graphs, the CGCG in which the connectivity radius depends on the type or colour or symbol or spin of the nodes. The empirical pair measure and the empirical locality measure are our main object of study.

Given a probability measure \( \nu \) on \( \Sigma_1 \) and a function \( r_n: \Sigma \times \Sigma \rightarrow (0, 1] \) we may define the randomly coloured random geometric graph or simply coloured random geometric graph \( G \) with \( n \) vertices as follows: Pick vertices \( x_1, \ldots, x_n \) at random independently according to the uniform distribution on \([0, 1]^d, d \in \mathbb{N}\). Assign to each vertex \( x_j \) colour \( \sigma(x_j) \) independently according to the colour law \( \nu \). Given the colours, we join any two vertices \( x_i, x_j (i \neq j) \) by an edge independently of everything else, if

\[
|x_i - x_j| \leq r_n[\sigma(x_i), \sigma(x_j)].
\]

In this article we shall refer to \( r_n(a,b) \), for \( a, b \in \Sigma \) as a connection radius, and always consider

\[
G = (\{(\sigma(x_i), \sigma(x_j)): i, j = 1, 2, 3, \ldots, n\}, E)
\]

under the joint law of graph and colour. We interpret \( G \) as coloured GRG with vertices \( x_1, \ldots, x_n \) chosen at random uniformly and independently from the vertices space \([0, 1]^2\). For the purposes of this study we restrict ourselves to the near intermediate cases, i.e. the connection radius \( r_n \) satisfies the condition \( nr_n^d(a,b) \rightarrow C_d(a,b) \) for all \( a, b \in \Sigma \), where \( C_d: \Sigma^2 \rightarrow [0, \infty) \) is a symmetric function, which is not identically equal to zero.

For any finite or countable set \( \Sigma \) we denote by \( \mathcal{P}(\Sigma) \) the space of probability measures, and by \( \tilde{\mathcal{P}}(\Sigma) \) the space of finite measures on \( \Sigma \), both endowed with the weak topology. By convention we write \( \mathbb{N} = \{0, 1, 2, \ldots\} \).

We associate with any coloured graph \( G \) a probability measure, the empirical colour measure \( \mathcal{L}^1_G \in \mathcal{P}(\Sigma) \), by

\[
\mathcal{L}^1_G(a) := \frac{1}{n} \sum_{j=1}^{n} \delta_{\sigma(x_j)}(a), \quad \text{for } a_1 \in \Sigma,
\]

and a symmetric finite measure, the empirical pair measure \( \mathcal{L}^2_G \in \tilde{\mathcal{P}}_n(\Sigma^2) \), by

\[
\mathcal{L}^2_G(a,b) := \frac{1}{n} \sum_{(ij) \in E} [\delta_{(\sigma(x_i),\sigma(x_j))} + \delta_{(\sigma(x_j),\sigma(x_i))}](a,b), \quad \text{for } (a,b) \in \Sigma^2.
\]
Note that the total mass the empirical pair measure is $2|E|/n$. Finally we define a further probability measure, the empirical neighbourhood measure $\mathcal{M}_G \in \mathcal{P}(\Sigma \times \mathbb{N})$, by

$$
\mathcal{M}_G(a, \ell) := \frac{1}{n} \sum_{j=1}^{n} \delta_{(\sigma(x_j), L(x_j))}(a, \ell), \quad \text{for } (a, \ell) \in \Sigma \times \mathbb{N},
$$

while $L(x_j) = (l^G(b), b \in \Sigma)$ and $l^G(b)$ is the number of vertices of colour $b$ connected to vertex $x_j$.

For any $\eta \in \mathcal{P}(\Sigma \times \mathbb{N}^\Sigma)$ we denote by $\eta_1$ the $\Sigma$-marginal of $\eta$ and for every $(b, a) \in \Sigma \times \Sigma$, let $\eta_2$ be the law of the pair $(a, l(b))$ under the measure $\eta$. Define the measure (finite), $(\eta(\cdot), \ell(\cdot)) \in \hat{\mathcal{P}}(\Sigma \times \Sigma)$ by

$$
\mathcal{H}_2(\eta)(b, a) := \sum_{l(b) \in \mathbb{N}} \eta_2(a, l(b)) l(b), \quad \text{for } a, b \in \Sigma
$$

and write $\mathcal{H}_1(\eta) = \eta_1$. We define the function $\mathcal{H}: \mathcal{P}(\Sigma \times \mathbb{N}^\Sigma) \to \mathcal{P}(\Sigma) \times \hat{\mathcal{P}}(\Sigma \times \Sigma)$ by $\mathcal{H}(\eta) = (\mathcal{H}_1(\eta), \mathcal{H}_2(\eta))$ and note that $\mathcal{H}(\mathcal{M}_G) = (L_2^G, L_2^G)$. Observe that $\mathcal{H}_1$ is a continuous function but $\mathcal{H}_2$ is discontinuous in the weak topology. In particular, in the summation $\sum_{l(b) \in \mathbb{N}} \eta_2(a, l(b)) l(b)$ the function $l(b)$ may be unbounded and so the functional $\eta \to \mathcal{H}_2(\eta)$ would not be continuous in the weak topology. We call a pair of measures $(\omega, \eta) \in \hat{\mathcal{P}}(\Sigma \times \Sigma) \times \mathcal{P}(\Sigma \times \mathbb{N}^\Sigma)$ sub-consistent if

$$
\mathcal{H}_2(\eta)(b, a) \leq \omega(b, a), \quad \text{for all } a, b \in \Sigma,
$$

and consistent if equality holds in (1). For a measure $\omega \in \hat{\mathcal{P}}_s(\Sigma^2)$ and a measure $\rho \in \mathcal{P}(\Sigma)$, we recall from (Doku-Amponsah and Mörters 2010) the rate function

$$
\mathcal{S}_1(\omega \parallel \rho) := H(\omega \parallel C_d \rho \otimes \rho) + \|C_d \rho \otimes \rho\| - \|\omega\|,
$$

where the measure $C_d \rho \otimes \rho \in \hat{\mathcal{P}}(\Sigma \times \Sigma)$ is defined by $C_d \rho \otimes \rho(a, b) = C_d(a, b) \rho(a) \rho(b)$ for $a, b \in \Sigma$. It is not hard to see that $\mathcal{S}_1(\omega \parallel \rho) \geq 0$ and equality holds if and only if $\omega = C_d \rho \otimes \rho$.

For every $(\omega, \eta) \in \hat{\mathcal{P}}_s(\Sigma \times \Sigma) \times \mathcal{P}(\Sigma \times \mathbb{N})$ define a probability measure $Q_{\text{pol}}^{(\omega, \eta)}$ on $\Sigma \times \mathbb{N}$ by

$$
Q_{\text{pol}}^{(\omega, \eta)}(a, \ell) := \eta_1(a) \prod_{b \in \Sigma} e^{-\omega(a,b)/\eta_1(a)} \frac{1}{\ell(b)!} \left(\frac{\omega(a,b)}{\eta_1(a)}\right)^{\ell(b)}, \quad \text{for } a \in \Sigma, \ell \in \mathbb{N}.
$$

We assume $d \in \mathbb{N}$ and write

$$
\Delta(d) = \begin{cases} 
\frac{d^{d/2}}{\Gamma\left(\frac{d+1}{2}\right)} & \text{if } d \geq 2 \\
1 & \text{if } d = 1,
\end{cases}
$$

where $\Gamma$ is the gamma function. We now state the principal theorem in this section the LDP for the empirical pair measure and the empirical locality measure.

**Theorem 1** Suppose that $G$ is a CRGG with colour law $\nu$ and connection radii $r_n: \Sigma \times \Sigma \to [0,1]$ satisfying $m_n^{d}(a,b) \to C_d(a,b)$ for some symmetric function $C_d$.
C: $\Sigma \times \Sigma \to [0, \infty)$ not identical to zero. Then, as $n \to \infty$, the pair $(\mathcal{L}_G^1, \mathcal{M}_G)$ satisfies an LDP in $\hat{\mathcal{P}}_s(\Sigma \times \Sigma) \times \mathcal{P}(\Sigma \times \mathbb{N})$ with good rate function

$$
J(\omega, \eta) = \begin{cases} 
H(\eta \parallel Q_{\text{pos}}^{(\omega, \eta)}) + H(\eta_1 \parallel \nu) + \frac{1}{2}J_2(\omega \parallel \eta_1) & \text{if } (\omega, \eta) \text{ consistent and } \eta_1 = \omega_2, \\
\infty & \text{otherwise}.
\end{cases}
$$

$$
J_2(\omega \parallel \eta_1) = \hat{J}_1(\omega \parallel \eta_1) - \|\omega\| \log \Delta(d) + (\Delta(d) - 1)\|C_d \eta_1 \otimes \eta_1\|.
$$

Remark 1 Note that the first three terms of the rate function is the same as the rate function of Doku-Amponsah and Mörters (2010, Theorem 2.1). Additionally, the extra term $\frac{1}{2}(-\|\omega\| \log \Delta(d) + (\Delta(d) - 1)\|C_d \eta_1 \otimes \eta_1\|)$ is positive and is as a result of the geometric $[0, 1]^d$ we have incorporated in the model. Moreover, on typical CGRG we have, $\eta_1 = \nu, \omega = \Delta(d)C \eta \otimes \eta_1$ and

$$
\eta(a, \ell) = \nu(a) \prod_{b \in \Sigma} e^{-\Delta(d)C(a, \ell)(\nu(b))} \frac{\Delta(d)C(a, \ell)(\nu(b))^{\ell(b)}}{\ell(b)!}, \quad \text{for all } (a, \ell) \in \Sigma \times \mathbb{N}.
$$

Hence, for some $\varepsilon$ we $\mathbb{P}\{|\mathcal{M}_G - \eta\| \geq \varepsilon\} \to 0$ as $n \to \infty$.

We write

$$
\lambda_1(\delta) := (\Delta(d) - 1) \frac{c}{2} - \frac{1}{2} \langle \delta \rangle \log \Delta(d)
$$

**Corollary 2** Suppose $D$ is the degree distribution of the random graph $\mathcal{G}(n, r_n)$, where the connectivity radius $r_n \in (0, 1)$ satisfies $nr_n^d \to c \in (0, \infty)$. Then, as $n \to \infty$, $D$ satisfies an LDP in the space $\mathcal{P}(\mathbb{N} \cup \{0\})$ with good rate function

$$
\lambda_2(\delta) = \begin{cases} 
H(d \parallel q(\delta)) + \frac{1}{2} \langle \delta \rangle \log \left(\frac{\langle \delta \rangle}{c}\right) - \frac{1}{2} \langle \delta \rangle + \frac{c}{2} & \text{if } \langle \delta \rangle < \infty, \\
\lambda_1(\delta) & \text{if } \langle \delta \rangle = \infty, 
\end{cases}
$$

where $q_k$ is a poisson distribution with parameter $k$, and $\langle \delta \rangle := \sum_{m=0}^{\infty} m \delta(m)$.

This rate function $\lambda_2$ compares very well with the rate function of Doku-Amponsah and Mörters (2010, Corollary 2.2) with the extra term $\lambda_1$ accounting for the geometric effect on the CGRG model.

Next we give a similar result as in O’Connell (1998), the LDP for the proportion of isolated vertices of the RGG.

$$
\xi_1(y) = (\Delta(d) - 1)cy(1 - y/2) + (1 - y) \left[ \log \left( \frac{1}{\Delta(d)} \right) - \frac{(\Delta(d) - 1)c(1 - y)}{2} \right]
$$

**Corollary 3** Suppose $D$ is the degree distribution of the random graph $\mathcal{G}(n, r_n)$, where the connectivity radius $r_n \in (0, 1)$ satisfies $nr_n^d \to c \in (0, \infty)$. Then, as $n \to \infty$, the proportion of isolated vertices, $D(0)$ satisfies an LDP in $[0, 1]$ with good rate function

$$
\xi_2(y) = y \log y + cy(1 - y/2) - (1 - y) \left[ \log \left( \frac{c}{a} \right) - \frac{(a - c(1 - y))^2}{2c(1 - y)} \right] + \xi_1(y),
$$

where $a = a(y, d)$ is the unique positive solution of $1 - e^{-a} = \frac{\Delta(d)c}{a}(1 - y)$.
From Corollary 3 we deduce that on a typical random geometric graphs the number of isolated vertices will grow like \( ne^{-\Delta(d)c} \). Thus, as \( n \to \infty \), the number of isolated vertices in the geometric random graphs converges to \( ne^{-\Delta(d)c} \) in probability. Again, the rate function \( \xi_2 \) above compares very well with the result of O’Connell (1998) with the extra term \( \xi_1 \) accounting for the influence of the geometric plane \([0, 1]^d\) on the model.

The joint LDP for the empirical colour measure and empirical pair measure of CGRG

**Theorem 4** Suppose that \( G \) is a CGRG with colour law \( v \) and connection radii \( r_n: \Sigma^2 \to [0, 1] \) satisfying \( nr_n^d(a, b) \to C_d(a, b) \) for some symmetric function \( C_d: \Sigma^2 \to [0, \infty) \) not identical to zero. Then, as \( n \to \infty \), the pair \( (\mathcal{L}_X^1, \mathcal{L}_X^2) \) satisfies an LDP in \( \mathcal{P}(\Sigma) \times \hat{\mathcal{P}}_v(\Sigma^2) \) with good rate function

\[
I(\eta_1, \omega) = H(\eta_1 \| v) + \frac{1}{2} \Omega_2(\omega \| \eta_1),
\]

where the measure \( C_{\eta_1} \otimes \eta_1 \in \hat{\mathcal{P}}_v(\Sigma \times \Sigma) \) is defined by \( C_{\eta_1} \otimes \eta_1(a, b) = C_d(a, b)\eta_1(a)\eta_1(b) \) for \( a, b \in \Sigma \).

Further, we state a corollary of Theorem 4 below.

**Corollary 5** Suppose that \( G \) is a CGRG graph with colour law \( v \) and connection radii \( r_n: \Sigma^2 \to [0, 1] \) satisfying \( nr_n^d(a, b) \to C_d(a, b) \) for some symmetric function \( C_d: \Sigma^2 \to [0, \infty) \) not identical to zero. Then, as \( n \to \infty \), the number of edges per vertex \( |E| / n \) of \( G \) satisfies an LDP in \([0, \infty)\) with good rate function

\[
\zeta(x) = x \log x - x + \inf_{y > 0} \left\{ \psi(y) - x \log(y) + y \right\},
\]

where \( \psi(y) = \inf H(\eta_1 \| v) \) over all probability vectors \( \eta_1 \) with \( \frac{1}{2} \Delta(d)\eta_1^T C\eta_1 = y \).

**Remark 2** By taking \( C_d(a, b) = C \) one will obtain \( \psi(y) = 0 \) for \( y = \frac{\Delta(d)c}{2} \), and \( \psi(y) = \infty \) otherwise, which establishes that \( |E| / n \) obeys an LDP in \([0, \infty)\) with good rate function

\[
\zeta(x) = x \log x - x + \inf_{y > 0} \left\{ \psi(y) - x \log \left( \frac{1}{2} y \right) + \frac{1}{2} y \right\},
\]

where \( \Delta(d)c = y \).

**Proof of Theorem 4**

**Change-of-measure**

For any two points \( U_1 \) and \( U_2 \) uniformly and independently chosen from the space \([0, 1]^d\) write

\[
F(t) := \mathbb{P}(|U_1 - U_2| \leq t).
\]

Further, given a function \( \tilde{f}: \Sigma \to \mathbb{R} \) and a symmetric function \( \tilde{g}: \Sigma^2 \to \mathbb{R} \), we define the constant \( U_f \) by

\[
U_f = \log \sum_{a \in \Sigma} e^{\tilde{f}(a)} v(a),
\]
and the function $\tilde{h}_n$: $\Sigma^2 \to \mathbb{R}$ by
\[
\tilde{h}_n(a, b) = \log \left[ \left( 1 - F(r_n(a, b)) + F(r_n(a, b))e^{\tilde{g}(a,b)} \right)^{-n} \right],
\]
for $a, b \in \Sigma$. We use $\tilde{f}$ and $\tilde{g}$ to define (for sufficiently large $n$) a new coloured random graph as follows:

- To the $n$ points $x_1, x_2, \ldots, x_n$ picked independently and uniformly in $[0, 1]^d$ we assign colours from $\Sigma$ independently and identically according to the colour law $\tilde{v}$ defined by
  \[
  \tilde{v}(a) = d^{(a-B)} v(a).
  \]
- Given any two points $x_u, x_v$ with $x_u$ carrying colour $a$ and $x_v$ carrying colour $b$, we connect vertex $x_u$ to vertex $x_v$ with probability
  \[
  F(\tilde{r}_n(a, b)) = \frac{F(r_n(a, b))e^{\tilde{g}(a,b)}}{1 - F(r_n(a, b)) + F(r_n(a, b))e^{\tilde{g}(a,b)}}.
  \]
We denote the transformed law by $\tilde{\mathbb{P}}$. We observe that $\tilde{v}$ is a probability measure and that $\tilde{\mathbb{P}}$ is absolutely continuous with respect to $\mathbb{P}$ as, for any coloured graph $G = ((\sigma(x)_j): j = 1, 2, 3, \ldots, n, E)$,
\[
\frac{d\tilde{\mathbb{P}}}{d\mathbb{P}}(G) = \prod_{u \in V} \frac{\tilde{v}(\sigma(x)_u)}{v(\sigma(x)_u)} \prod_{(u,v) \in E} \frac{F(\tilde{r}_n(\sigma(x)_u), \sigma(x)_v))}{F(r_n(\sigma(x)_u), \sigma(x)_v))} \prod_{(u,v) \notin E} \frac{1 - F(\tilde{r}_n(\sigma(x)_u), \sigma(x)_v))}{1 - F(r_n(\sigma(x)_u), \sigma(x)_v))} \\
= \prod_{u \in V} \frac{\tilde{v}(\sigma(x)_u)}{v(\sigma(x)_u)} \prod_{(u,v) \in E} \frac{F(\tilde{r}_n(\sigma(x)_u), \sigma(x)_v))}{F(r_n(\sigma(x)_u), \sigma(x)_v))} \times \prod_{(u,v) \notin E} \frac{n - nF(\tilde{r}_n(\sigma(x)_u), \sigma(x)_v))}{n - nF(r_n(\sigma(x)_u), \sigma(x)_v))} \\
= \prod_{u \in V} e^{\tilde{f}(\sigma(x)_u) - U_f} \prod_{(u,v) \in E} e^{\tilde{g}(\sigma(x)_u), \sigma(x)_v))} \prod_{(u,v) \notin E} e^{\tilde{h}_n(\sigma(x)_u), \sigma(x)_v))} \\
= \exp \left( n \left\langle \mathcal{L}_G^1, \tilde{f} - U_f \right\rangle + n \left\langle \frac{1}{2} \mathcal{L}_G^2, \tilde{g} \right\rangle + n \left\langle \frac{1}{2} \mathcal{L}_G^1 \otimes \mathcal{L}_G^1, \tilde{h}_n \right\rangle - \left\langle \frac{1}{2} \mathcal{L}_\Delta^1, \tilde{h}_n \right\rangle \right),
\]
where
\[
L_\Delta^1 = \frac{1}{n} \sum_{u \in V} \delta(\sigma(x)_u, \sigma(x)_u).
\]
We write $(g, \omega) := \sum_{a,b \in \Sigma} g(a, b)\omega(a, b)$ for $\omega \in \tilde{\mathbb{P}}(\Sigma^2)$, and $(f, \rho) := \sum_{a \in \Sigma} f(a)\rho(a)$ for $\rho \in \mathcal{P}(\Sigma)$, and note that
\[
F(r_n(a, b)) = \Delta(d)r_n^d(a, b), \quad \text{for all } a, b \in \Sigma^2.
\]
i.e. the volume of a $d$-dimensional (hyper)sphere with radius $r(a, b)$ satisfying $nr_n^d(a, b) \to C_d(a, b)$.

The following lemmas will be useful in the proofs of main Lemmas.

**Lemma 1**  (Euler’s lemma) If $nr_n^d(a, b) \to C_d(a, b)$ for every $a, b \in \Sigma$, then

$$
\lim_{n \to \infty} \left[ 1 + \frac{\alpha \Delta(d) C_d(a, b)}{\alpha F(r_n(a, b))} \right]^n = e^{\alpha \Delta(d) C_d(a, b)}, \quad \text{for all } a, b \in \Sigma \text{ and } \alpha \in \mathbb{R}.
$$

**Proof**  Observe that, for any $\varepsilon > 0$ and for large $n$ we have

$$
\left[ 1 + \frac{\alpha \Delta(d) C_d(a, b) - \varepsilon}{n} \right]^n \leq \left[ 1 + \frac{\alpha \Delta(d) C_d(a, b)}{n} \right]^n \leq \left[ 1 + \frac{\alpha \Delta(d) C_d(a, b) + \varepsilon}{n} \right]^n,
$$

by the point-wise convergence. Hence by the sandwich theorem and Euler’s formula we get (6).

We write

$$
P(n)(\omega) := \mathbb{P}\left\{ L_{\mathcal{G}}^1 = \omega \right\}.
$$

**Lemma 2**  The family of measures $(P_n: n \in \mathbb{N})$ is exponentially tight on $\mathcal{P}(\Sigma)$.

**Proof**  We use coupling argument, see the proof of Doku-Amponsah and Mörters (2010, Lemma 5.1) to show that, for every $\theta > 0$, there exists $N \in \mathbb{N}$ such that

$$
\limsup_{n \to \infty} \frac{1}{n} \mathbb{P}\{|E| > nN\} \leq -\theta.
$$

To begin, let $c(d) > \max_{a,b \in \Sigma} C_d(a, b) > 0$ and $nr_n^d(a, b) \to c(d)$. Using similar coupling arguments as in see the proof of Doku-Amponsah and Mörters (2010, Lemma 5.1), we can define, for all sufficiently large $n$, a coloured random graph $\tilde{X}$ with vertices $x_1, \ldots, x_n$ chosen uniformly from the vertices space $[0, 1]^d$, colour law $\eta$ and connectivity probability $p_n = \mathbb{P}\{ \|x_i - x_j\| \leq r_n(c) \} = \Delta(d) r_n^d$, for all $i \neq j$ such that any edge present in $\mathcal{G}$ is also present in $\tilde{X}$. Let $|\tilde{E}|$ be the number of edges of $\tilde{X}$. Using the binomial formula and Euler’s formula, we have that

$$
\mathbb{P}\{|\tilde{E}| \geq nl\} \leq e^{-nl} \mathbb{E}\left[ e^{\tilde{E}} \right] = e^{-nl} \sum_{k=0}^{n(n-1)/2} \left( \begin{array}{c} n(n-1)/2 \\ k \end{array} \right) \left( p_n \right)^k \left( 1 - p_n \right)^{n(n-1)/2-k}
$$

$$
= e^{-nl} (1 - p_n + np_n)^{n(n-1)/2} \leq e^{-nl} e^{\eta c \Delta(d) (\eta - 1 + o(1))},
$$

where we used $np_n = \Delta(d) nr_n^d \to \Delta(d) c$ in the last step. Now given $\theta > 0$ choose $N \in \mathbb{N}$ such that $N > \theta + \Delta(d) c (\eta - 1)$ and observe that, for sufficiently large $n$,

$$
\mathbb{P}\{|E| \geq nN\} \leq \mathbb{P}\{|\tilde{E}| \geq nN\} \leq e^{-n\theta},
$$

which implies the statement.

**Proof of the upper bound in Theorem 4**

We denote by $C_1$ the space of functions on $\Sigma$ and by $C_2$ the space of symmetric functions on $\Sigma^2$, and define
\[ \hat{I}(\eta_1, \omega) = \sup_{f \in C_1, g \in C_2} \left\{ \sum_{a \in \Sigma} (f(a) - U_f) \eta_1(a) + \frac{1}{2} \sum_{a, b \in \Sigma} g(a, b) \omega(a, b) + \frac{\Delta(d)}{2} \sum_{a, b \in \Sigma} (1 - \epsilon g(a, b)) C_d(a, b) \eta_1(a) \eta_1(b) \right\} \]

for \((\eta_1, \omega) \in \mathcal{P}(\Sigma) \times \mathcal{P}(\Sigma^2)\)

**Lemma 3**  For each closed set \(G \subset \mathcal{P}(\Sigma) \times \hat{\mathcal{P}}_\omega(\Sigma^2)\), we have

\[ \limsup_{n \to \infty} \frac{1}{n} \log \mathbb{P}\left\{ \left( \mathcal{L}^2_{\hat{\omega}}, \mathcal{L}^2_{\hat{\mu}} \right) \in F \right\} \leq - \inf_{(\eta_1, \omega) \in F} \hat{I}(\eta_1, \omega). \]

**Proof**  First let \(\tilde{f} \in C_1\) and \(\tilde{g} \in C_2\) be arbitrary. Define \(\tilde{\beta}: \Sigma^2 \to \mathbb{R}\) by

\[ \tilde{\beta}(a, b) = \Delta(d)(1 - \epsilon g(a, b)) C_d(a, b). \]

Observe that, by Lemma 1, \(\tilde{\beta}(a, b) = \lim_{n \to \infty} \tilde{h}_n(a, b)\) for all \(a, b \in \Sigma\), recalling the definition of \(\tilde{h}_n\) from (4). Hence, by (5), for sufficiently large \(n\),

\[ e^{\max_{a \in \Sigma} \beta(a, a)} \geq e^{\frac{1}{2} \tilde{L} \tilde{h}_n} = \mathbb{E}\left\{ e^{n(\mathcal{L}^2_{\hat{\omega}} \tilde{f} - U_f) + n(\frac{1}{2} \mathcal{L}^2_{\hat{\mu}} \tilde{g}) + n(\frac{1}{2} \mathcal{L}^2_{\hat{\mu}} \tilde{h}_n)} \right\}, \]

where \(\tilde{L} = \frac{1}{n} \sum_{a \in \Sigma} \delta(\sigma(x_{\eta_1}), x_{\eta_1}))\) and therefore,

\[ \limsup_{n \to \infty} \frac{1}{n} \log \mathbb{E}\left\{ e^{n(\mathcal{L}^2_{\hat{\omega}} \tilde{f} - U_f) + n(\frac{1}{2} \mathcal{L}^2_{\hat{\mu}} \tilde{g}) + n(\frac{1}{2} \mathcal{L}^2_{\hat{\mu}} \tilde{h}_n)} \right\} \leq 0. \] (7)

Given \(\varepsilon > 0\) let \(\hat{I}_\varepsilon(\eta_1, \omega) = \min\{\hat{I}(\eta_1, \omega), \varepsilon^{-1}\} - \varepsilon\). Suppose that \((\eta_1, \omega) \in G\) and observe that \(\hat{I}(\eta_1, \omega) > \hat{I}_\varepsilon(\eta_1, \omega)\). We now fix \(\tilde{f} \in C_1\) and \(\tilde{g} \in C_2\) such that

\[ \langle \tilde{f} - U_f, \eta_1 \rangle + \frac{1}{2} \langle \tilde{g}, \omega \rangle + \frac{1}{2} \langle \tilde{\beta}, \eta_1 \otimes \eta_1 \rangle \geq \hat{I}_\varepsilon(\eta_1, \omega). \]

As \(\Sigma\) is finite, there exist open neighbourhoods \(B^1_{\eta_1}\) and \(B^2_{\omega}\) of \(\eta_1, \omega\) such that

\[ \inf_{\eta_1 \in B^1_{\eta_1}} \inf_{\omega \in B^2_{\omega}} \left\{ \langle \tilde{f} - U_f, \eta_1 \rangle + \frac{1}{2} \langle \tilde{g}, \omega \rangle + \frac{1}{2} \langle \tilde{\beta}, \eta_1 \otimes \eta_1 \rangle \right\} \geq \hat{I}_\varepsilon(\eta_1, \omega) - \varepsilon. \]

Using Chebyshev’s inequality and (7) we have that

\[ \limsup_{n \to \infty} \frac{1}{n} \log \mathbb{P}\left\{ \left( \mathcal{L}^2_{\hat{\omega}}, \mathcal{L}^2_{\hat{\mu}} \right) \in B^1_{\eta_1} \times B^2_{\omega} \right\} \leq - \hat{I}_\varepsilon(\eta_1, \omega) + \varepsilon \]

for \(\varepsilon > 0\). Therefore, by choosing \(N(\varepsilon) \in \mathbb{N}\) such that

\[ \lim \sup_{n \to \infty} \frac{1}{n} \log \mathbb{P}\{|E| > nN(\varepsilon)| \} \leq -\varepsilon^{-1}. \] (9)
For this \( N(\varepsilon) \), define the set \( K_{N(\varepsilon)} \) by

\[
K_{N(\varepsilon)} = \left\{ (\eta_1, \omega) \in \mathcal{P}(\Sigma) \times \tilde{P}_e(\Sigma^2) : \| \omega \| \leq 2N(\varepsilon) \right\},
\]

and recall that \( \| \mathcal{L}_g^2 \| = 2|E|/n \). The set \( K_{N(\varepsilon)} \cap F \) is compact and therefore may be covered by finitely many sets \( B_{\eta_1}^1 \times B_{\omega_0}^2, r = 1, \ldots, m \) with \((\eta_1, \omega_r) \in F \) for \( r = 1, \ldots, m \). Consequently,

\[
\mathbb{P}\left\{ \mathcal{L}_g^1, \mathcal{L}_g^2 \in F \right\} \leq \sum_{r=1}^{m} \mathbb{P}\left\{ (L_{g}^1, L_{g}^2) \in B_{\eta_1}^1 \times B_{\omega_0}^2 \right\} + \mathbb{P}\left\{ (L_{g}^1, L_{g}^2) \notin K_{N(\varepsilon)} \right\}.
\]

We may now use (8) and (9) to obtain, for all sufficiently small \( \varepsilon > 0 \),

\[
\limsup_{n \to \infty} \frac{1}{n} \log \mathbb{P}\left\{ (L_{g}^1, L_{g}^2) \in F \right\} \leq \max_{r=1}^{m} \left( \limsup_{n \to \infty} \frac{1}{n} \log \mathbb{P}\left\{ (L_{g}^1, L_{g}^2) \in B_{\eta_1}^1 \times B_{\omega_0}^2 \right\} \right) \vee (-\varepsilon)^{-1} \leq \left( -\inf_{(\eta_1, \omega) \in G} \hat{I}(\eta_1, \omega) + \varepsilon \right) \vee (-\varepsilon)^{-1}.
\]

Taking \( \varepsilon \downarrow 0 \) we get the desired statement. \( \square \)

Next, we express the rate function in term of relative entropies, see for example Dembo and Zeitouni (1998, 2.15), and consequently show that it is a good rate function. Recall the definition of the function \( \hat{I} \) from Theorem 4.

**Lemma 4**

(i) \( \hat{I}(\eta_1, \omega) = I(\eta_1, \omega) \), for any \((\eta_1, \omega) \in \mathcal{P}(\Sigma) \times \tilde{P}_e(\Sigma^2)\).

(ii) \( \hat{I}(\eta_1, \omega) \) is a good rate function and

(iii) \( \mathcal{H}_2(\omega \| \eta_1) \geq 0 \) with equality if and only if \( \omega = \Delta(d)C_d\eta_1 \otimes \eta_1 \).

**Proof**

(i) Suppose that \( \omega \ll \Delta(d)C_d\eta_1 \otimes \eta_1 \). Then, there exists \( a_0, b_0 \in \Sigma \) with \( C\eta_1 \otimes \eta_1 (a_0, b_0) = 0 \) and \( \omega(a_0, b_0) > 0 \). Define \( \hat{g} : \Sigma^2 \to \mathbb{R} \) by

\[
\hat{g}(a, b) = \log \left[ K(1_{(a_0, b_0)}(a, b) + 1_{(b_0, a_0)}(a, b)) + 1 \right], \quad \text{for } a, b \in \Sigma \text{ and } K > 0.
\]

For this choice of \( \hat{g} \) and \( f = 0 \) we have

\[
\sum_{a \in \Sigma} (f(a) - U_f)\eta_1(a) + \sum_{a, b \in \Sigma} \frac{1}{2}\hat{g}(a, b)\omega(a, b) + \sum_{a, b \in \Sigma} \frac{\Delta(d)}{2} (1 - e^{\hat{g}(a, b)}) C_d(a, b) \eta_1(a)\eta_1(b)
\]

\[
\geq \frac{\Delta(d)}{2} \log(K + 1)\omega(a_0, b_0) \to \infty, \quad \text{for } K \uparrow \infty.
\]

Now suppose that \( \omega \ll C\eta_1 \otimes \eta_1 \). We have

\[
\hat{I}(\eta_1, \omega) = \sup_{f \in C_1} \left\{ \sum_{a \in \Sigma} \left( f(a) - \log \sum_{a \in \Sigma} e^{f(a)} \nu(a) \right) \eta_1(a) \right\}
\]

\[
+ \frac{\Delta(d)}{2} \sum_{a, b \in \Sigma} C_d(a, b) \eta_1(a)\eta_1(b)
\]

\[
+ \frac{1}{2} \sup_{g \in C_2} \left\{ \sum_{a, b \in \Sigma} g(a, b)\omega(a, b) - \Delta(d) \sum_{a, b \in \Sigma} e^{g(a, b)} C_d(a, b) \eta_1(a)\eta_1(b) \right\}.
\]

By the variational characterization of relative entropy, the first term equals \( H(\eta_1 \| \nu) \). By the substitution \( h = \Delta(d)e^{\hat{g}} \frac{C_d\eta_1 \otimes \eta_1}{\omega} \) the last term equals...
\[
\sup_{c \leq d} \sum_{a,b \in \Sigma} \left[ \log \left( \frac{\omega(a,b)}{\Delta(d)C_d(a,b)\eta_1(a)\eta_1(b)} \right) - h(a,b) \right] \omega(a,b) = \\
\sup_{c \leq d} \sum_{a,b \in \Sigma} (\log h(a,b) - h(a,b)) \omega(a,b) + \sum_{a,b \in \Sigma} \log \left( \frac{\omega(a,b)}{\Delta(d)C_d(a,b)\eta_1(a)\eta_1(b)} \right) \omega(a,b) = \\
-\|\omega\| + H(\omega \| \Delta(d)C_d\eta_1 \otimes \eta_1),
\]

where we have used \( \sup_{x>0} \log x - x = -1 \) in the last step. This yields that
\[
\hat{I}(\eta_1, \omega) = I(\eta_1, \omega).
\]

(ii) Recall from (3) and the definition of \( \hat{F}_2 \) that \( I(\eta_1, \omega) = H(\omega \| \nu) + \frac{1}{2} \Delta(d)C_d\eta_1 \otimes \eta_1 + \|\Delta(d)C_d\eta_1 \otimes \eta_1\| - \frac{1}{2} \|\omega\|. \) All summands are continuous in \( \eta_1, \omega \) and thus \( I \) is a rate function. Moreover, for all \( \alpha < \infty \), the level sets \( \{ \| I \| \leq \alpha \} \) are contained in the bounded set \( \{ (\eta_1, \omega) \in \mathcal{P}(\Sigma) \times \tilde{\mathcal{P}}_\nu(\Sigma^2) : \hat{F}_2(\omega \| \eta_1) \leq \alpha \} \) and are therefore compact. Consequently, \( I \) is a good rate function.

(iii) Consider the nonnegative function \( \xi(x) = x \log x - x + 1 \), for \( x > 0 \), \( \xi(0) = 1 \), which has its only root in \( x = 1 \). Note that
\[
\hat{F}_2(\omega \| \eta_1) = \left\{ \begin{array} \left\{ \int \xi \circ g \; d(\Delta(d)C_d\omega \otimes \omega) \\ \infty \\
\end{array} \right. \\
\text{if } g := \frac{d\omega}{d(\Delta(d)C_d\eta_1 \otimes \eta_1)} \geq 0 \text{ exists,} \\
\text{otherwise.}
\]
Hence \( \hat{F}_2(\omega \| \eta_1) \geq 0 \), and if \( \omega = \Delta(d)C_d\eta_1 \otimes \eta_1 \), then \( \int \xi \circ g \; d(\Delta(d)C_d\omega \otimes \omega) = \xi(1) = 0 \) and so \( \hat{F}_2(\Delta(d)C_d\eta_1 \otimes \eta_1 \| \omega) = 0 \). Conversely, if \( \hat{F}_2(\omega \| \omega) = 0 \), then \( \omega(a,b) > 0 \) implies \( C_d\eta_1 \otimes \eta_1(a,b) > 0 \), which then implies \( \xi \circ g(a,b) = 0 \) and further \( g(a,b) = 1 \). Hence \( \omega = \Delta(d)C_d\eta_1 \otimes \eta_1 \), which completes the proof of (iii).

\[\Box\]

**Proof of the lower bound in Theorem 4**

We obtain the lower bound of Theorem 4 from the upper bound as follows:

**Lemma 5**  **For every open set** \( O \subset \mathcal{P}(\Sigma) \times \tilde{\mathcal{P}}_\nu(\Sigma^2) \), we have
\[
\liminf_{n \to \infty} \frac{1}{n} \log \mathbb{P}\left\{ \left( \mathcal{L}_\omega^1, \mathcal{L}_\omega^2 \right) \in O \right\} \geq - \inf_{(\eta_1, \omega) \in O} I(\eta_1, \omega).
\]

**Proof** Suppose \( (\eta_1, \omega) \in O \), with \( \omega \ll \Delta(d)C_d\eta_1 \otimes \eta_1 \). Define \( \tilde{f}_\omega : \Sigma \to \mathbb{R} \) by
\[
\tilde{f}_\omega(a) = \left\{ \begin{array} \left\{ \log \frac{\omega(a,b)}{\Delta(d)C_d(a,b)\eta_1(a)\eta_1(b)} \\ 0, \end{array} \right. \\
\text{if } \eta_1(a) > 0, \\
\text{otherwise.}
\]
and \( \tilde{g}_\omega : \Sigma^2 \to \mathbb{R} \) by
\[
\tilde{g}_\omega(a,b) = \left\{ \begin{array} \left\{ \log \frac{\omega(a,b)}{\Delta(d)C_d(a,b)\eta_1(a)\eta_1(b)} \\ 0, \end{array} \right. \\
\text{if } \omega(a,b) > 0, \\
\text{otherwise.}
\]
In addition, we let \( \tilde{\beta}_\omega(a,b) = \Delta(d)C_d(a,b)(1 - \tilde{g}_\omega(a,b)) \) and note that \( \tilde{\beta}_\omega(a,b) = \lim_{n \to \infty} \tilde{h}_{\omega,n}(a,b) \), for all \( a, b \in \Sigma \) where
\[
\tilde{h}_{\omega,n}(a,b) = \log \left[ \left( 1 - F(r_n(a,b)) + F(r_n(a,b))\tilde{g}_\omega(a,b) \right)^{-n} \right].
\]
Choose $B^1_{\eta_1}, B^2_{\omega}$ open neighbourhoods of $\eta_1, \omega$, such that $B^1_{\eta_1}, x B^2_{\omega} \subset O$ and for all $(\tilde{\omega}, \tilde{\omega}) \in B^1_{\eta_1} \times B^2_{\omega}$

$$\tilde{f}_{\omega}(\eta_1) + \frac{1}{2} \langle \tilde{g}_{\omega}, \omega \rangle + \frac{1}{2} \langle \tilde{\beta}_{\omega}, \eta_1 \otimes \eta_1 \rangle - \varepsilon \leq \tilde{f}_{\omega}(\tilde{\eta}_1) + \frac{1}{2} \langle \tilde{g}_{\omega}, \tilde{\omega} \rangle + \frac{1}{2} \langle \tilde{\beta}_{\omega}, \tilde{\eta}_1 \otimes \tilde{\eta}_1 \rangle.$$ 

We now use $\tilde{P}$, the probability measure obtained by transforming $P$ using the functions $\tilde{f}_{\omega}$, $\tilde{g}_{\omega}$. Note that the colour law in the transformed measure is now $\eta_1$, and the connectivity radii $\tilde{r}_n(a, b)$ satisfy

$$n \tilde{r}_n(a, b) \to \omega(a, b)/(\eta_1(a) \eta_1(b)) =: \tilde{C}_d(a, b), \quad as \ n \to \infty.$$ 

Using (5), we obtain

$$\mathbb{P}\left\{ \left( L^1_{G_{\omega}}, L^2_{G_{\omega}} \right) \in O \right\} \geq \mathbb{P}\left\{ \left\{ \left( L^1_{G_{\omega}}, L^2_{G_{\omega}} \right) \in B^1_{\eta_1} \times B^2_{\omega} \right\} \right\}$$

$$= \mathbb{P}\left\{ \prod_{u \in V} e^{-\tilde{f}_{\omega}(\sigma(u))} \prod_{(u,v) \in E} e^{-\tilde{g}_{\omega}(\sigma(u), \sigma(v))} \prod_{u \in V} e^{-\frac{1}{2} \tilde{\beta}_{\omega}(\sigma(u))} \right\}$$

$$\geq \exp\left( -n \tilde{f}_{\omega}(\omega) - n \frac{1}{2} \langle \tilde{g}_{\omega}, \omega \rangle - n \frac{1}{2} \langle \tilde{\beta}_{\omega}, \eta_1 \otimes \eta_1 \rangle + m + nE \right) \times P\left\{ \left( L^1_{G_{\omega}}, L^2_{G_{\omega}} \right) \in B^1_{\eta_1} \times B^2_{\omega} \right\},$$

where $m := 0 \land \min_{x \in \chi} \tilde{\beta}(a, a)$. Therefore, by (6), we have

$$\liminf_{\eta_1 \to \infty} \frac{1}{n} \log \mathbb{P}\left\{ \left( L^1_{G_{\omega}}, L^2_{G_{\omega}} \right) \in O \right\}$$

$$\geq -n \tilde{f}_{\omega}(\omega) - n \frac{1}{2} \langle \tilde{g}_{\omega}, \omega \rangle - n \frac{1}{2} \langle \tilde{\beta}_{\omega}, \eta_1 \otimes \eta_1 \rangle - \varepsilon + \liminf_{\eta_1 \to \infty} \frac{1}{n} \log \tilde{P}\left\{ \left( L^1_{G_{\omega}}, L^2_{G_{\omega}} \right) \in B^1_{\eta_1} \times B^2_{\omega} \right\}. $$

The result follows once we prove that

$$\liminf_{n \to \infty} \frac{1}{n} \log \tilde{P}\left\{ \left( L^1_{G_{\omega}}, L^2_{G_{\omega}} \right) \in B^1_{\eta_1} \times B^2_{\omega} \right\} = 0. \quad (11)$$

We use the upper bound (but now with the law $\tilde{P}$ replaced by $\tilde{P}$) to prove (11). Then we obtain

$$\limsup_{n \to \infty} \frac{1}{n} \log \tilde{P}\left\{ \left( L^1_{G_{\omega}}, L^2_{G_{\omega}} \right) \in B^1_{\eta_1} \times B^2_{\omega} \right\} \leq - \inf_{(\tilde{\rho}, \tilde{\omega}) \in \tilde{F}} \tilde{I}(\tilde{\rho}, \tilde{\omega}),$$

where $\tilde{F} := (B^1_{\eta_1} \times B^2_{\omega})^C$ and $\tilde{I}(\tilde{\rho}, \tilde{\omega}) := H(\tilde{\omega} \parallel \omega) + \frac{1}{2} \tilde{S}_2(\tilde{\omega} \parallel \tilde{\rho})$. It therefore suffices to show that the infimum is positive. Suppose for contradiction that there exists a sequence $(\tilde{\rho}_n, \tilde{\omega}_n) \in \tilde{F}$ with $I(\tilde{\rho}_n, \tilde{\omega}_n) \downarrow 0$. Then, because $I$ is a good rate function and its level sets are compact, and by lower semi-continuity of the mapping $(\tilde{\rho}, \tilde{\omega}) \mapsto I(\tilde{\rho}, \tilde{\omega})$, we can construct a limit point $(\tilde{\rho}, \tilde{\omega}) \in \tilde{F}$ with $I(\tilde{\rho}, \tilde{\omega}) = 0$. By Lemma 4 this implies $H(\tilde{\omega} \parallel \eta_1) = 0$ and $S_2(\tilde{\omega} \parallel \eta_1) = 0$, hence $\tilde{\rho} = \eta_1$, and $\tilde{\omega} = \tilde{C}_d \eta_1 \otimes \eta_1 = \omega$ contradicting $(\tilde{\rho}, \tilde{\omega}) \in \tilde{F}$. 

\textbf{Proof of Theorem 1}

For any $n \in \mathbb{N}$ we define
\[ \mathcal{P}_n(\Sigma) := \{ \rho \in \mathcal{P}(\Sigma) : n \rho(a) \in \mathbb{N} \text{ for all } a \in \Sigma \}, \]
\[ \hat{\mathcal{P}}_n(\Sigma \times \Sigma) := \left\{ \omega \in \hat{\mathcal{P}}_s(\Sigma \times \Sigma) : \frac{n}{1 + 1} \omega(a, b) \in \mathbb{N} \text{ for all } a, b \in \Sigma \right\}. \]

We denote by \( \Theta_n := \mathcal{P}_n(\Sigma) \times \hat{\mathcal{P}}_n(\Sigma \times \Sigma) \) and \( \Theta := \mathcal{P}(\Sigma) \times \hat{\mathcal{P}}_s(\Sigma \times \Sigma) \). With

\[ P_{(\rho_n, \omega_n)}^{(n)}(\eta_n) := \mathbb{P}\{ \mathcal{M}_G = \eta_n \mid \mathcal{H}(\mathcal{M}_G) = (\rho_n, \omega_n) \}, \]
\[ P^{(n)}(\rho_n, \omega_n) := \mathbb{P}\{ (L^1_G, L^2_G) = (\rho_n, \omega_n) \} \]

the joint distribution of \( L^1_G, L^2_G \) and \( \mathcal{M}_G \) is the mixture of \( P_{(\rho_n, \omega_n)}^{(n)} \) with \( P^{(n)}(\rho_n, \omega_n) \) defined as

\[ d\hat{\mathcal{P}}^{(n)}(\rho_n, \omega_n, \eta_n) := dP_{(\rho_n, \omega_n)}^{(n)}(\eta_n) dP^{(n)}(\rho_n, \omega_n). \] (12)

Biggins (2004, Theorem 5(b)) gives criteria for the validity of large deviation principles for the mixtures and for the goodness of the rate function if individual large deviation principles are known. The following three lemmas ensure validity of these conditions.

We recall from Lemma 6 that the family of measures \((\mathcal{P}_n : n \in \mathbb{N})\) is exponentially tight on \( \Theta \times \mathcal{P}(\Sigma \times \Sigma) \).

**Lemma 6** (Doku-Amponsah and Mörters 2010) The family of measures \((\hat{\mathcal{P}}_n : n \in \mathbb{N})\) is exponentially tight on \( \Theta \times \mathcal{P}(\Sigma \times \Sigma) \).

Define the function

\[ \hat{J} : \Theta \times \mathcal{P}(\Sigma \times \Sigma) \to [0, \infty], \quad \hat{J}((\eta_1, \omega), \eta) = \hat{J}_{(\eta_1, \omega)}(\eta), \]

where

\[ \hat{J}_{(\eta_1, \omega)}(\eta) = \left\{ \begin{array}{ll} H(\eta \parallel Q^{(\omega, \eta)}_{\text{poi}}) & \text{if } (\omega, \eta) \text{ is consistent and } \eta_1 = \omega_2 \\ \infty & \text{otherwise.} \end{array} \right. \] (13)

**Lemma 7** (Doku-Amponsah and Mörters 2010) \( \hat{J} \) is lower semi-continuous.

By Biggins (2004, Theorem 5(b)) the two previous lemmas and the large deviation principles we have established Theorem 2.2 and Doku-Amponsah (2015, Theorem 2.1) ensure that under \((\hat{\mathcal{P}}_n)\) the random variables \((\rho_n, \omega_n, \eta_n)\) satisfy a large deviation principle on \( \mathcal{P}(\Sigma) \times \hat{\mathcal{P}}_s(\Sigma \times \Sigma) \times \mathcal{P}(\Sigma \times \mathbb{N}) \) with good rate function

\[ \hat{j}(\eta_1, \omega, \eta) = \left\{ \begin{array}{ll} H(\eta_1 \parallel v) + \frac{1}{2} S_2(\omega \parallel \Sigma) + H(\eta \parallel Q^{(\omega, \eta)}_{\text{poi}}), & \text{if } (\omega, \eta) \text{ is consistent and } \eta_1 = \omega_2, \\ \infty, & \text{otherwise.} \end{array} \right. \]

By projection onto the last two components we obtain the large deviation principle as stated in Theorem 1 from the contraction principle, see e.g. Dembo et al. (2005, Theorem 4.2.1).
Proof of Corollaries 2, 3, and 5

We derive the theorems from Theorem 1 by applying the contraction principle, see e.g. Dembo and Zeitouni (1998, Theorem 4.2.1). In fact Theorem 1 and the contraction principle imply a large deviation principle for $D$. It just remains to simplify the rate functions.

Proof of Theorem 2

Note that, in the case of an uncoloured RGG graphs, the function $C$ degenerates to a constant $c$, $L^2_0 = |E|/n \in [0, \infty)$ and $\mathcal{M}_C = D \in \mathcal{P}(\mathbb{N} \cup \{0\})$. Theorem 1 and the contraction principle imply a large deviation principle for $D$ with good rate function $\lambda_2(\delta) = \inf \{H(x, \delta): x \geq 0\}$

$$\lambda_2(\delta) = \inf \{H(x, \delta): x \geq 0\} = \inf \left\{ x \log x - \frac{1}{2} x \log \Delta(d)c + \frac{1}{2} \Delta(d)c - \frac{1}{2} x: \langle \delta \rangle \leq x \right\},$$

which is to be understood as infinity if $\langle d \rangle$ is infinite. We denote by $\lambda^x(\delta)$ the expression inside the infimum. For any $\epsilon > 0$, we have

$$\lambda^{\epsilon}(\delta) - \lambda^0(\delta) = \epsilon + \frac{\langle \delta \rangle - \epsilon}{2} \log \frac{\langle \delta \rangle}{\Delta(d)c} \geq \epsilon + \frac{\langle \delta \rangle - \epsilon}{2} \left( -\frac{\langle \delta \rangle}{\Delta(d)c} \right) + \frac{\epsilon}{2} \log \frac{\langle \delta \rangle}{\Delta(d)c} > 0,$$

so that the minimum is attained at $x = \Delta(d)\langle \delta \rangle$.

Proof of Corollary 3

Corollary 3 follows from Theorem 2 and the contraction principle applied to the continuous linear map $G: \mathcal{P}(\mathbb{N} \cup \{0\}) \to [0, 1]$ defined by $G(\delta) = \delta(0)$. Thus, Theorem 2 implies the large deviation principle for $G(D) = W$ with the good rate function $\xi_2(y) = \inf \{\lambda_2(\delta): \delta(0) = y, \langle \delta \rangle < \infty\}$. We recall the definition of $\lambda_2(\delta)$ and observe that $\xi_2(y)$ can be expressed as

$$\xi_2(y) = \inf_{b \geq 0} \inf_{d \in \mathcal{P}(\mathbb{N} \cup \{0\}), \delta(0) = y, \langle \delta \rangle = \langle \delta \rangle} \left\{ \frac{1}{2} c + y \log y + \frac{b^2}{2\Delta(d)c} + \sum_{k=1}^{\infty} \delta(k) \log \frac{\delta(k)}{q_b(k)} - b(1 - y) \right\}.$$

Now, using Jensen's inequality, we have that

$$\sum_{k=1}^{\infty} \delta(k) \log \frac{\delta(k)}{q_b(k)} \geq (1 - y) \log \frac{(1 - y)}{(1 - e^{-b})},$$

with equality if $\delta(k) = \frac{(1 - y)}{(1 - e^{-b})} q_b(k)$, for all $k \in \mathbb{N}$. Therefore, we have the inequality

$$\inf \{\lambda_2(\delta): \delta(0)$$

$$= y, \langle \delta \rangle < \infty\} \geq \inf \left\{ \frac{1}{2} c + y \log y + \frac{b^2}{2\Delta(d)c} + (1 - y) \log \frac{(1 - y)}{(1 - e^{-b})} - b(1 - y): b \geq 0 \right\}.$$
\[ b \mapsto \frac{1}{2} \Delta(d)c + y \log y + \frac{b^2}{2\Delta(d)c} + (1 - y) \log \frac{(1-y)}{(1-e^{-b})} - b(1-y) \text{ on } (0, \infty) \text{ is attained at the value } b = a, \text{ where } a \text{ is the positive solution of our equation. We obtain the form of } \xi \text{ in Corollary 3 by observing that}
\]
\[
\frac{a(y)^2 + (\Delta(d)c)^2 - 2\Delta(d)c a(y)(1-y)}{2\Delta(d)c} = \frac{\Delta(d)c y}{2}(2-y) + \frac{1}{2\Delta(d)c} (a(y) - \Delta(d)c(1-y))^2.
\]

**Proof of Corollary 5**

We define the continuous linear map \( W: \mathcal{P}(\Sigma) \times \mathcal{P}_*^2(\Sigma^2) \to [0, \infty) \) by \( W(\eta_1, \omega) = \frac{1}{2} \| \omega \| \), and infer from Theorem 4 and the contraction principle that \( W(L^2_G, L^2_B) = |E|/n \) satisfies a large deviation principle in \([0, \infty)\) with the good rate function

\[
\zeta(y) = \inf \{ I(\eta_1, \omega): W(\eta_1, \omega) = y \}.
\]

To obtain the form of the rate in the corollary, the infimum is reformulated as unconstrained optimization problem (by normalising \( \omega \))

\[
\inf_{\eta_1 \in \mathcal{P}_*^2(\Sigma)} \{ H(\eta_1 \| \nu) + yH(\omega \| \Delta(d)C\eta_1 \otimes \eta_1) + y \log 2y + \frac{\Delta(d)}{2} \| C\omega \otimes \omega \| - y \}. \quad (15)
\]

By Jensen’s inequality \( H(\omega \| \Delta(d)C\eta_1 \otimes \eta_1) \geq - \log \| \Delta(d)C\eta_1 \otimes \eta_1 \| \), with equality if \( \omega = \frac{C\eta_1 \otimes \eta_1}{\| C\eta_1 \otimes \eta_1 \|} \), and hence, by symmetry of \( C \) we have

\[
\min_{\omega \in \mathcal{P}_*^2(\Sigma)} \{ H(\eta_1 \| \nu) + yH(\omega \| \Delta(d)C\eta_1 \otimes \eta_1) + y \log 2y + \frac{\Delta(d)}{2} \| C\eta_1 \otimes \eta_1 \| - y \}
\]
\[
= H(\eta_1 \| \nu) - y \log \| \Delta(d)C\eta_1 \otimes \eta_1 \| + y \log 2y + \frac{\Delta(d)}{2} \| C\eta_1 \otimes \eta_1 \| - y.
\]

The form given in Corollary 5 follows by defining

\[
y = \frac{1}{2} \Delta(d) \sum_{a,b \in \Sigma} C_d(a, b) \eta_1(a) \eta_1(b).
\]

**Conclusion**

In this work, we have proved joint large deviation principle for the empirical pair measure and empirical locality measure of the near intermediate CGRG models. From this result we have obtained asymptotic results about useful graph quantities such as number of edges per vertex, the degree distribution and the proportion of isolated vertices for the near intermediate CGRG models. The rate functions of all these large deviation principles compared very well with the rate functions of the results for coloured random graph models by Doku-Amponsah and Mörters (2010), with some extra terms accounting for the geometric effect in the CGRG models. An important future research direction is to formulate and prove an Asymptotic Equipartition Property for Networked Data Structures Modelled as the CGRG, and then a possible Coding or Approximate Pattern Matching Algorithms for such Networks. One could also investigate the Statistical Mechanics on the CGRG.
Acknowledgements
This extension has been mentioned in the author’s PhD Thesis at University of Bath.

Competing interests
The author declares that he has no competing interests.

Received: 16 January 2016   Accepted: 29 June 2016
Published online: 20 July 2016

References
Biggins JD (2004) Large deviations for mixtures. Electron Commun Probab 9:60–71
Bordenave C, Caputo P (2015) Large deviations of empirical neighbor-hood distribution in sparse random graphs. Probab
Theory Relat Fields 163(1–2):149–222
Biggins JD, Penman DB (2009) Large deviations in randomly coloured random graphs. Electron Comm Probab
14:290–301
Cannings C, Penman DB (2003) Handbook of statistics 21. Stochastic processes: modelling and simulation. In: Shanbhag
DN, Rao CR (eds) Models of random graphs and their applications. Elsevier, Amsterdam, pp 51–91
Dembo A, Zeitouni O (1998) Large deviations techniques and applications. Springer, New York
Dembo A, Mörters P, Sheffield S (2005) Large deviations of Markov chains indexed by random trees. Ann Inst Henri
Poincaré Probab Stat 41:971–996
Doku-Amponsah K (2006) Large deviations and basic information theory for hierarchical and networked data structures.
PhD thesis, Bath
Doku-Amponsah K (2014) Exponential approximation, method of types for empirical neighbourhood measures of
random graphs by random allocation. Int J Stat Probab 3(2):110–120
Doku-Amponsah K (2015) Large deviation result for the empirical locality measure of typed random geometric graphs.
Int J Stat Prob 4(1):94–99
Doku-Amponsah K (2016) Large deviation, Basic information theory for wireless sensor networks. Presented at College of
Basic and Applied Sciences, Science Development Platform 2016, University of Ghana. arxiv:1512.08050
Doku-Amponsah K, Mörters P (2010) Large deviation principles for empirical measures of coloured random graphs. Ann
Appl Probab 20(6):1989–2021. doi:10.1214/09-AAP647
Mukherjee S (2014) Large deviation for the empirical degree distribution of an Erdos–Renyi graph. arxiv:1310.4160
O’Connell N (1998) Some large deviation results for sparse random graphs. Probab Theory Relat Fields 110:277–285
Penrose MD (2003) Random geometric graphs. Oxford University Press, Oxford

Submit your manuscript to a SpringerOpen journal and benefit from:
► Convenient online submission
► Rigorous peer review
► Immediate publication on acceptance
► Open access: articles freely available online
► High visibility within the field
► Retaining the copyright to your article

Submit your next manuscript at ► springeropen.com