Impact of Neighborhood Socioeconomic Conditions on the Risk of Stroke in Japan

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ABSTRACT

Background: Neighborhood deprivation has been shown in many studies to be an influential factor in cardiovascular disease risk. However, no previous studies have examined the effect of neighborhood socioeconomic conditions on the risk of stroke in Asian countries.

Methods: This study investigated whether neighborhood deprivation was associated with the risk of stroke and stroke death using data from the Japan Public Health Center-based Prospective Study. We calculated the adjusted hazard ratios of stroke mortality (mean follow-up, 16.4 years) and stroke incidence (mean follow-up, 15.4 years) according to the area deprivation index (ADI) among 90,843 Japanese men and women aged 40–69 years. A Cox proportional-hazard regression model using a shared frailty model was applied.

Results: The adjusted hazard ratios of stroke incidence, in order of increasing deprivation with reference to the least deprived area, were 1.16 (95% CI, 1.04–1.29), 1.12 (95% CI, 1.00–1.26), 1.18 (95% CI, 1.02–1.35), and 1.19 (95% CI, 1.01–1.41), after adjustment for individual socioeconomic conditions. Behavioral and psychosocial factors attenuated the association, but the association remained significant. The associations were explained by adjusting for biological cardiovascular risk factors. No significant association with stroke mortality was identified.

Conclusions: Our results indicate that the neighborhood deprivation level influences stroke incidence in Japan, suggesting that area socioeconomic conditions could be a potential target for public health intervention to reduce the risk of stroke.

Key words: neighborhood; stroke; socioeconomic status; poverty areas; Japan

INTRODUCTION

A wealth of evidence has consistently demonstrated that the level of neighborhood deprivation is an influential factor in the risk of cardiovascular disease.1 There is a strong, independent association between living in a socioeconomically disadvantaged neighborhood and risk of coronary heart disease,2,3 total stroke,4,5 and ischemic stroke,6,7 even after adjustment for individual socioeconomic position. However, to our knowledge, there have been no studies examining neighborhood socioeconomic conditions in relation to the risk of stroke among general residents in Asian countries.

Evidence of social inequalities in cardiovascular disease has been accumulating in Japan11; several studies have shown differences in mortality, morbidity, and risk factors for cardiovascular disease—stroke in particular—according to indicators of individual socioeconomic conditions.12–16 Regional inequalities in health have also been repeatedly reported. Several ecological studies have indicated that municipal-level socioeconomic conditions exert an influence
on the health of the population in Japan: lower municipal-level socioeconomic conditions have had an adverse influence on population health, health behaviors (eg, smoking, excess alcohol consumption, and physical inactivity), and body mass index (BMI). All of those studies were cross-sectional or ecological studies. One exception, a recently-conducted multi-level study, indicated that higher municipal deprivation levels were associated with increased risk of cardiovascular mortality among men. However, thus far, no prospective studies have been conducted to examine the impact of finer neighborhood-level socioeconomic conditions on individual risk of developing cardiovascular diseases.

In this study, we aimed to investigate whether census-based neighborhood-level socioeconomic conditions are associated with stroke mortality and incidence, using data from a large prospective cohort study carried out in a Japanese non-metropolitan setting.

METHODS

Study cohort
We used data from cohorts 1 and 2 of the Japan Public Health Center-based Prospective Study (JPHC Study), a large population-based prospective study of 140,420 men and women aged 40–69 years. Two Public Health Center (PHC) areas in metropolitan Tokyo and Osaka were excluded from the present analysis (n = 23,524) because no data on stroke incidence were available. Of the remaining 116,896 participants, 9 were excluded as ineligible (7 were non-Japanese and 2 had moved prior to the start of the study).

The baseline self-administered questionnaire was distributed to all registered participants in 1990 (for cohort 1) and 1993 (for cohort 2), and the overall response rate was 81.6%. Of the 95,405 participants, we excluded 32 ineligible subjects and 3650 with a history of cancer or cardiovascular diseases. The remaining 91,723 participants were considered eligible for inclusion in this analysis.

Measurement
To examine the effect of neighborhood deprivation on stroke mortality and incidence, this study used area deprivation data at the level of the chocho-aza (CA) unit, the smallest administrative unit in Japan, which is roughly comparable to a European parish or a U.S. block-group-size neighborhood. We excluded four subjects whose addresses could not be geocoded. We calculated the CA-level area deprivation index (ADI), which is a composite indicator consisting of the weighted sum of a number of census-based variables. The ADI was derived and has been described in detail by Nakaya. In brief, we constructed an ADI for each of 695 CAs. The ADI was calculated using deprivation-related census-based variables (eg, the proportions of elderly couple households, elderly single households, single-mother households, sales and service workers, agricultural workers, blue-collar workers, and non-employed persons) in the 1995 population census in each area unit. The ADI was tested against ecological datasets of all-cause and various cancer mortalities at the municipal level across Japan and showed consistent positive relationships with most mortality indices.

In addition, we confirmed that higher ADI was associated with increased risk of all-cause death using the same method. We considered occupation as an indicator of individual socioeconomic conditions. Occupation was categorized into nine groups: professional, management, office work, sales/service, transportation/communication industries and not classifiable, manual labor, agriculture/forestry/fishery, non-working, and information missing.

Age, gender, PHC district, and population density at the CA level were considered as confounding factors, and biological cardiovascular risk factors (hypertension, diabetes or hyperlipidemia, and overweight) and behavioral and psychosocial factors (perceived psychological stress, marital status, smoking, alcohol intake, and physical activity) were considered as hypothesized mediating factors.

Study participants
Of the 91,723 eligible participants in this cohort study, we further excluded 880 subjects because no census information was provided by the statistical bureau for their CAs, or their CAs had a small number of households (less than 25). The remaining 90,843 men and women were included in the analysis. The median area was 2.39 km² (range, 0.02–56.28 km²), median population was 1234 (range, 68–9330), and median number of households was 363 (range, 25–3254). The study was approved by the human ethics review committee of the National Cancer Center.

Confirmation of stroke incidence and mortality
A total of 78 hospitals were registered within the administrative districts of the JPHC cohort. All were major hospitals capable of treating patients with stroke. Physicians blinded to the patients’ lifestyle data reviewed the medical records at each hospital. Strokes were confirmed according to the criteria of the National Survey of Stroke, which requires the presence of a focal neurological deficit of sudden or rapid onset lasting at least 24 hours or until death. All of the registered hospitals were major hospitals with admission facilities for acute cardiovascular events and were equipped with a computed tomographic scanning and/or magnetic resonance imaging apparatus. In addition to annual surveillance, we conducted another search for unidentified events. When subjects reported a history of nonfatal stroke on the 5- and 10-year follow-up questionnaire and had not been registered as stroke cases, we asked by letter or telephone about the onset of stroke and for permission to review medical records. Thus, we assume that most of the acute stroke cases were captured.
All death certificates were forwarded centrally to the Ministry of Health, Labour and Welfare, and coded for the National Vital Statistics. Registration of death is required by the Family Registration Law and is believed to be complete in Japan. The underlying causes of deaths were defined according to the International Classification of Diseases, 10th Revision, as deaths from stroke (codes I60 to I69).

Stroke deaths and events were included in the analyses if they occurred after the date of return of the baseline questionnaire and before January 1, 2010 (for cohort 1), and January 1, 2008 (for cohort 2).

Statistical analysis
The outcomes for this study were defined as newly occurring stroke incidence and deaths during the study period. Person-years were counted from the date of the return of the baseline survey until one of the following end points. For the analysis of stroke incidence, person-years were censored at the date of disease diagnosis, the date of emigration from the study area, the date of death, or the end of the study period (December 31, 2009, for cohort 1 and December 31, 2007, for cohort 2), whichever came first. For the analysis of stroke deaths, person-years were censored at the date of emigration from the study area, the date of death, or the end of the study period (December 31, 2010, for cohorts 1 and 2), whichever came first. For persons who were lost to follow-up, the last confirmed date of their presence in the study area was used as the date of censoring.

We estimated adjusted hazard ratios (HRs) and 95% confidence intervals (CIs) of stroke incidence and mortality by applying a Cox proportional-hazard regression model using a shared frailty model that allows individuals to be nested within neighborhoods and the intercept to vary between neighborhoods.\(^2^7\) We assumed two area levels: broad regional differences in living situations (adjusted for by including dummy variables of PHC district for each person) and the CA level (covering possible clustering tendencies within the same neighborhood according to the random intercept of the shared frailty function). Model fitting procedures were carried out using the PHREG command with RANDOM statement by SAS ver. 9.3 (SAS Institute Inc., Cary, NC, USA).

The dummy variables of PHC district and the random effect were included for all of the fitted models to calculate adjusted HRs and 95% CIs (Model 1). Further adjustment was made for occupation category considered as an individual socioeconomic indicator, as well as age, gender, and population density considered as potential confounding variables (Model 2). To examine the mediating mechanism, social and behavioral factors (marital status, perceived psychological stress, smoking behavior, frequency and amount of alcohol intake, and frequency of physical activity during leisure; Model 3) and biological cardiovascular risk factors (overweight, medical history of hypertension, diabetes, or hyperlipidemia; Model 4) were inserted into the model.

All analyses were conducted with SAS statistical package version 9.3 (SAS Institute Inc.).

RESULTS
Table 1 lists the characteristics of the study population and their associations with ADI level. Participants residing in the most deprived neighborhoods were more likely to be older, to have agriculture/forestry/fishery jobs, to be overweight, and to live alone, and were less likely to smoke, to have a medical history of hypertension, and to perceive they had psychological stress, compared with those in less deprived neighborhoods. We identified 1147 stroke deaths during a mean follow-up period of 16.4 years, while 4410 stroke events were documented during a mean follow-up period of 15.4 years.

Table 2 shows the adjusted HRs and 95% CIs of stroke mortality and incidence according to ADI level. A significant association was identified between neighborhood deprivation level and the risk of developing stroke. The adjusted HRs for risk of developing stroke, in order of increasing deprivation with reference to the least deprived area, were 1.16 (95% CI, 1.04–1.29), 1.12 (95% CI, 1.00–1.26), 1.18 (95% CI, 1.02–1.35), and 1.19 (95% CI, 1.01–1.41), after adjustment for individual socioeconomic conditions. Behavioral and psychosocial factors attenuated the association, but the association remained significant. The association was explained by adjusting for biological cardiovascular risk factors. On the other hand, no association was identified between neighborhood deprivation level and the risk of stroke death.

DISCUSSION
In this cohort, we observed that neighborhood deprivation level was associated with total stroke incidence after adjustment for individual socioeconomic indicators, and living in more deprived neighborhoods increased the risk of developing stroke regardless of individual socioeconomic conditions. The association remained after adjustment for individual socioeconomic indicators, which suggests that the neighborhood socioeconomic conditions have an impact on disparities in stroke incidence in Japanese society. Our results suggest that neighborhood socioeconomic conditions could be a potential target for public health interventions to reduce the risk of stroke in certain areas.

Our results were consistent with results of previous studies in Western societies that found that residing in disadvantaged neighborhoods is associated with increased risk of stroke.\(^1^,\(^8^–\(^1^0\)\)

Neighborhood socioeconomic conditions are assumed to be associated with the physical and social environments of neighborhoods, which impact cardiovascular health mediated through cardiovascular risk factors. For example, there is a higher prevalence of health-damaging behaviors and
### Table 1. Characteristics of study subjects

| Variable                          | Total          | Area Deprivation Index | P-value for difference<sup>a</sup> |
|-----------------------------------|----------------|------------------------|------------------------------------|
|                                  | 90,843         | 18,159, 17,992, 18,230, 18,282, 18,180 |
| Age, years (Mean (SD))            | 50.3 (7.6)     | 49.8 (7.9), 49.9 (7.4), 49.9 (7.4), 50.5 (7.5), 51.3 (7.7) | <0.0001 |
| BMI, kg/m² (Mean (SD))            | 23.5 (3.1)     | 23.2 (2.9), 23.2 (2.9), 23.3 (3.0), 23.8 (3.1), 24.0 (3.2) | <0.0001 |
| Gender: Men                       | 43,337 (48)    | 8827 (49), 8701 (48), 8599 (47), 8583 (47), 8627 (47) | 0.003   |
| Gender: Occupation                |                |                        |                                    |
| Professional                      | 3587 (4)       | 747 (4), 738 (4), 791 (4), 667 (4), 644 (4) | <0.0001 |
| Manager                           | 2898 (3)       | 915 (5), 640 (4), 555 (3), 435 (2), 353 (2) |          |
| Clerk                             | 7227 (8)       | 1654 (9), 1640 (9), 1602 (9), 1242 (7), 1089 (6) |          |
| Agriculture/Forestry/Fishery      | 21,790 (24)    | 3358 (18), 3418 (19), 3817 (21), 5335 (29), 5862 (32) |          |
| Service/Sales                     | 7187 (8)       | 1717 (9), 1405 (8), 1440 (8), 1452 (8), 1173 (6) |          |
| Manual job                        | 10,573 (12)    | 2022 (11), 2181 (12), 2451 (13), 2083 (11), 1836 (10) |          |
| No job                            | 18,596 (20)    | 3802 (21), 4379 (24), 4058 (22), 3395 (19), 2962 (16) |          |
| Security/transportation/communication/others | 17,759 (20) | 3742 (21), 3350 (19), 3208 (18), 3428 (19), 4031 (22) |          |
| Marital status: Married           | 73,356 (84)    | 14,809 (82), 14,666 (80), 14,226 (78), 13,420 (79) | <0.0001 |
| Marital status: Unmarried         | 16,761 (16)    | 3047 (18), 3399 (19), 3893 (21), 3766 (21) |          |
| Marital status: Missing           | 726 (1)        | 136 (1), 165 (1), 163 (1), 94 (1) |          |
| Perceived psychological stress: Low| 14,195 (16)    | 2449 (14), 2626 (14), 3091 (17), 3514 (19) | <0.0001 |
| Perceived psychological stress: Medium| 57,233 (63)  | 11426 (63), 11389 (63), 141661 (64), 11518 (63) |          |
| Perceived psychological stress: High| 17,923 (20)   | 3954 (22), 3867 (21), 3270 (18), 2971 (16) |          |
| Smoking: Never smoker             | 54,344 (60)    | 10,245 (57), 10,838 (59), 11,415 (62), 11,586 (64) | <0.0001 |
| Smoking: Quitter                  | 10,507 (12)    | 2117 (12), 2044 (11), 1957 (11), 2158 (12) |          |
| Smoking: Current smoker           | 25,563 (28)    | 5507 (31), 5287 (29), 4852 (27), 4369 (24) |          |
| Smoking: Missing                  | 429 (0.5)      | 135 (0.7), 108 (0.6), 61 (0.3), 58 (0.3), 67 (0.4) |          |
| Ethanol intake: No                | 46,815 (52)    | 8708 (48), 8743 (49), 8902 (48), 9988 (54), 10,474 (57) | <0.0001 |
| Ethanol intake: Occasional drinker| 8628 (10)      | 1583 (9), 1681 (9), 1762 (10), 1848 (10), 1754 (10) |          |
| Ethanol intake: 149 g and less per week | 10,337 (11) | 2122 (12), 2341 (13), 1859 (10), 1663 (9) |          |
| Ethanol intake: 150-299 g per week | 8494 (10)      | 1836 (10), 1817 (10), 1613 (9), 1362 (8) |          |
| Ethanol intake: 300-449 g per week | 7331 (9)       | 1635 (9), 942 (10), 792 (5), 708 (4) |          |
| Ethanol intake: 450 g and more per week | 7331 (8)  | 1508 (8), 1501 (8), 1405 (8), 1429 (8) |          |
| Physical activity at leisure time: No | 2836 (3)     | 540 (3), 453 (3), 522 (3), 531 (3), 790 (4) |          |
| Physical activity at leisure time: 1–3 times per month | 39,268 (43)  | 9030 (50), 7311 (41), 7512 (41), 7266 (40), 8149 (45) | <0.0001 |
| Physical activity at leisure time: 1–2 times per week | 8398 (9)      | 1887 (10), 1768 (10), 1618 (9), 1536 (8), 1589 (9) |          |
| Physical activity at leisure time: 3 times and more per week | 12,492 (14)  | 2651 (15), 2265 (13), 2412 (13), 2702 (15), 2462 (14) |          |
| Hypertension: Yes                 | 16,575 (18)    | 3398 (19), 3306 (18), 3259 (18), 3083 (17) | <0.0001 |
| Hypertension: Missing             | 141 (0.2)      | 47 (0.3), 21 (0.1), 17 (0.1), 23 (0.1) |          |
| Diabetes mellitus: Yes            | 4210 (5)       | 840 (5), 867 (5), 883 (5), 767 (4) | 0.0001   |
| Diabetes mellitus: Missing        | 141 (0.2)      | 47 (0.3), 22 (0.1), 16 (0.1), 23 (0.1) |          |
| Hyperlipidemia: Yes               | 1530 (2)       | 423 (2), 347 (2), 326 (2), 251 (1), 183 (1) | <0.0001 |
| Overweight (BMI ≥ 25)             | 25,630 (28)    | 4394 (24), 4422 (25), 4851 (27), 5717 (31), 6246 (34) | <0.0001 |
| Overweight (BMI ≥ 25): Missing    | 1029 (1)       | 175 (1), 172 (1), 233 (1), 232 (1), 217 (1) |          |

BMI, body mass index; SD, standard deviation.
<sup>a</sup>ANOVA test or Chi-square test were used.

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Table 2. Multivariate hazard ratios and 95% confidence intervals of stroke incidence and mortality according to area deprivation level among 90,843 men and women

| Area Deprivation Index | 0 (Least deprived) | 1 | 2 | 3 | 4 (Most deprived) | Trend P-value | Log-likelihood | Random effects (SD) |
|------------------------|--------------------|----|----|----|-------------------|---------------|-----------------|---------------------|
| **Death**              |                    |    |    |    |                   |               |                 |                     |
| Number of cases        | 288,735            | 300,655 | 304,956 | 302,767 | 296,639         |               |                 |                     |
| Person-years           | 202                | 225 | 242 | 226 | 252              |               |                 |                     |
| Model 0                | 1.00               | 1.17 (0.95, 1.43) | 1.32 (1.06, 1.65) | 1.36 (1.05, 1.75) | 1.79 (1.34, 2.39) | 0.004          | −12.858          | 0.08 (0.03)         |
| Model 1                | 1.00               | 1.00 (0.82, 1.22) | 1.04 (0.84, 1.28) | 0.99 (0.76, 1.27) | 1.02 (0.76, 1.38) | 0.76           | −12.495          | 0.001 (0.019)       |
| Model 2                | 1.00               | 1.00 (0.82, 1.22) | 1.04 (0.84, 1.29) | 0.99 (0.76, 1.28) | 1.02 (0.76, 1.37) | 0.75           | −12.448          | 0.001 (0.019)       |
| Model 3                | 1.00               | 1.00 (0.82, 1.22) | 1.03 (0.83, 1.29) | 0.99 (0.77, 1.28) | 1.02 (0.76, 1.38) | 0.75           | −12.450          | 0.002 (0.019)       |
| Model 4                | 1.00               | 1.06 (0.96, 1.18) | 1.02 (0.91, 1.14) | 1.06 (0.93, 1.22) | 1.05 (0.90, 1.23) | 0.13           | −47.984          | 0.003 (0.005)       |
| **Incidence**          |                    |    |    |    |                   |               |                 |                     |
| Number of cases        | 735                | 911 | 879 | 959 | 926              |               |                 |                     |
| Person-years           | 273,578            | 282,021 | 283,995 | 282,766 | 274,064         |               |                 |                     |
| Model 0                | 1.00               | 1.20 (1.06, 1.35) | 1.14 (1.00, 1.30) | 1.24 (1.07, 1.44) | 1.32 (1.12, 1.56) | 0.01           | −49.061          | 0.049 (0.010)       |
| Model 1                | 1.00               | 1.16 (1.04, 1.29) | 1.12 (1.00, 1.26) | 1.18 (1.02, 1.35) | 1.19 (1.01, 1.41) | 0.13           | −48.209          | 0.008 (0.006)       |
| Model 2                | 1.00               | 1.13 (1.02, 1.26) | 1.09 (0.97, 1.23) | 1.14 (1.00, 1.31) | 1.13 (0.96, 1.33) | 0.14           | −48.094          | 0.008 (0.006)       |
| Model 3                | 1.00               | 1.06 (0.96, 1.18) | 1.02 (0.91, 1.14) | 1.06 (0.93, 1.22) | 1.05 (0.90, 1.23) | 0.13           | −47.984          | 0.003 (0.005)       |

SD, standard deviation.
Model 0 = Null model.
Model 1 = Area Deprivation Index (ADI) and PHC district were in the model.
Model 2 = Model 1 + age, gender, occupation and population density.
Model 3 = Model 2 + Behavioral and psychosocial factors (smoking, ethanol intake, physical activity, perceived psychological stress, marital status).
Model 4 = Model 2 + CVD biological risk factors (overweight, medical history of hypertension/diabetes/hyperlipidemia).

Reduced use of healthcare resources in disadvantaged neighborhoods.8,28-32

Our mediation analysis showed some attenuation of HRs after adjusting for behavioral and psychosocial risk factors at an individual level, but the association between ADI and stroke risk remained significant. However, residual confounding from other important risk factors, such as mental health (eg, depression) or measurement errors in our variables, could influence the association. Biological cardiovascular risk factors explained the association between ADI and stroke risk, which suggests that the impact of neighborhood deprivation level on stroke risk could be a result of unequal distribution of biological cardiovascular risk factors. Further detailed research is needed to examine how the neighborhood socioeconomic factors influence stroke incidence in Japanese society.

We identified no association of ADI with risk of death from stroke, which was inconsistent with studies in Western societies.33 One possible explanation for the discrepancy could be an inappropriate level of analysis of stroke mortality. In terms of stroke death, secondary prevention, such as acute treatment and medical services, could be crucial. In Japan, provision of medical services has been implemented over large areas, such as medical districts. Thus, our level of analysis may be inappropriate to detect an area effect on stroke mortality.

Several limitations of this study warrant discussion. First, the generalizability of our study results could be limited. The data included nine public health districts in a non-metropolitan setting, so we should be cautious in interpreting our results, especially in metropolitan areas. Second, there was a 5-year gap between census data used for the ADI and the baseline data. However, the area socioeconomic characteristics were not likely to change dramatically in Japan, especially in non-metropolitan areas, so we assumed that the 5-year gap did not have a major effect on our results. Third, we used CA units as an indirect proxy for neighborhood, but we cannot negate the possibility that they may not reflect meaningful neighborhood boundaries relevant to cardiovascular disease. Fourth, we adjusted for individual-level socioeconomic conditions using occupation only; however, occupation may not be a strong enough indicator for individual socioeconomic conditions. Unfortunately, we did not have sufficient information on the participants’ socioeconomic status to fully adjust for individual socioeconomic conditions.

In summary, we found that residing in a deprived neighborhood influences the risk of stroke in Japan. The results indicate that neighborhood deprivation contributes to disparities in the risk of stroke in Japanese society, and suggest that related socioeconomic factors could be a potential target for public health intervention to reduce stroke incidence. Future research is necessary to further investigate how neighborhood physical and social factors influence cardiovascular risk.

ONLINE ONLY MATERIAL

Abstract in Japanese.
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