Stochastic output feedback MPC with intermittent observations

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Abstract

This paper designs a model predictive control (MPC) law for constrained linear systems with stochastic additive disturbances and noisy measurements, minimizing a discounted cost subject to a discounted expectation constraint. It is assumed that sensor data is lost with a known probability. Taking into account the data losses modelled by a Bernoulli process, we parameterise the predicted control policy as an affine function of future observations and obtain a convex linear-quadratic optimal control problem. Constraint satisfaction and a discounted cost bound are ensured without imposing bounds on the distributions of the disturbance and noise inputs. In addition, the average long-run undiscounted closed loop cost is shown to be finite if the discount factor takes appropriate values. We analyse robustness of the proposed control law with respect to possible uncertainties in the arrival probability of sensor data and we bound the impact of these uncertainties on constraint satisfaction and the discounted cost. Numerical simulations are provided to illustrate these results.

Key words: Model predictive control, output feedback, packet drops, chance constraints, convex optimisation.

1 Introduction

Applications of wireless sensor networks are becoming omnipresent, for example in healthcare monitoring, weather forecasting, autonomous driving and building management systems. These networks can be time-varying and subject to congestion. Networked control systems that rely on sensor signals transmitted over communication channels must tolerate communication delays and data losses. Such measurement uncertainties pose additional challenges for estimator and controller design, especially when constraints are present.

A large body of work exists on control problems involving estimation with intermittent observations. This is largely concerned with the Random Riccati Equation (Wang and Guo, 1999), and the main issue addressed is boundedness of the state estimation error covariance. In Sinopoli et al. (2004), packet losses are modelled by an independent and identically distributed (i.i.d.) Bernoulli process and fundamental results are reported, such as the existence of and bounds on the critical value for the arrival probability of the observation update and convergence properties of the algebraic Riccati equation for Kalman filters with intermittent observations. This is the basis of further work in Kar et al. (2012) showing that the estimation error covariance sequence $\Sigma_0, \Sigma_1, \ldots$ converges in distribution to a unique invariant distribution with finite mean when the observation arrival probability exceeds the critical value. In Shi et al. (2010), the boundedness of the estimation error covariance is studied from a probabilistic perspective and bounds on the probability of $\Sigma_k \preceq \Sigma$ (henceforth denoted $\mathbb{P}\{\Sigma_k \preceq \Sigma\}$) are derived. A similar metric is used in Mo and Sinopoli (2012) to study the boundedness of estimation error covariance. This work considers the tail distribution of the trace of the estimation error covariance and the decay rate of $\mathbb{P}\{\text{tr}(\Sigma_k) > \text{tr}(\Sigma)\}$ (where $\text{tr}(\cdot)$ denotes the trace of a matrix) with packet losses being modelled by a Markov process. For the case of Markovian packet losses, it is possible to account for temporal correlation of communication channel conditions, but the stability analysis is more involved without an i.i.d. assumption (You et al., 2011; Huang and Dey, 2007). In contrast to the aforementioned work focus-
ing on limiting behaviours of estimation error covariance. Huang and Bitmead (2013) provide a different perspective by analysing transient properties. This work considers an intermittent quantised communication link and characterises the cumulative probability distribution function of the escape time at which system states or state estimates first exit a given compact set.

The results mentioned above do not easily carry over to constrained cases, and problems involving constrained estimation and control under intermittent and noisy observations have received less attention. In Ren et al. (2013), the problem of minimising an average estimation error covariance over a finite horizon is studied subject to an energy constraint on the remote sensor. Leong and Devi (2012) consider a similar problem subject to an average transmission power constraint, while the probability of packet losses at each time instant is conditional on transmission power. In Mishra et al. (2020), the problem of controlling stable linear systems with Gaussian disturbances and measurement noise is considered subject to hard input constraints, where both sensor and control channels are unreliable. This work assumes that remote sensors equipped with computing capabilities send state estimates instead of raw sensor measurements, which could be costly and unreasonable in certain situations.

Assuming sensor measurements will be lost with a known probability according to a Bernoulli process, this paper proposes an MPC algorithm with state estimation to minimise a discounted cost subject to a discounted expectation constraint. The system model has additive disturbances and measurement noise with probability distributions that are not assumed to be known exactly and may have infinite support. The discount setting is common to many control problems (e.g. Bertsekas, 1995; Blackwell, 1965; Van Parys et al., 2013; Kouvaritakis et al., 2003; Kamgarpour and Summers, 2017; Feinberg and Shwartz, 1999), and an appropriate discount factor can provide stability guarantees (Postovan et al., 2017). In financial engineering applications and economics, discounting is typically used to cost future activities at the present time. More generally, the effects of uncertain predictions can be mitigated by discounting controller objectives and constraints, reflecting the considerations that near-term benefits are more important than potential future gains and that current actions should not imperil possible future development (Kouvaritakis et al., 2006). Examples include sustainable development (Kouvaritakis et al., 2006), public policy analysis (Dasgupta, 2008) and ecosystem management (Clark, 1973) where different discount factors can affect harvesting policies and species populations. Relative to our earlier work in Yan et al. (2020), this paper includes additional results on the terminal term in the cost, on stability in the discount setting and on robustness analysis with respect to uncertainties in the arrival probability of sensor data, plus more extensive numerical studies. The main contributions of this paper are summarised as follows:

- The discount factors ensure both the cost function and constraint are well defined without bounds on the disturbance and measurement noise, and achieve a trade-off between short-term and long-term performance;
- We parameterise predicted control inputs as affine functions of future output measurements and show that the problem of optimising the associated feedback gains is convex. This allows the distributions of future states to be controlled explicitly even when output measurements are lost;
- We derive a bound on the discounted cost and ensure constraint satisfaction for the closed loop system;
- We provide two different conditions for ensuring stochastic stability in closed loop;
- We investigate the impact of uncertainty in the arrival probability of sensor data on constraint satisfaction and on the discounted cost.

The paper is organised as follows. The control problem is described in Section 2, and the controller parameterisation and implementation are introduced in Section 3. Section 4 derives the expressions of predicted state and control sequences via their first and second moments. Using these expressions, we formulate the MPC optimisation problem in Section 5. Section 6 derives a closed loop discounted cost bound and analyses constraint satisfaction. The averaged undiscounted closed loop cost accumulated over an infinite horizon is shown to be finite and an explicit bound is given in Section 7. We conduct a robustness analysis with respect to possible uncertainties in the arrival probability of sensor measurements in Section 8. Sections 9 and 10 provide a numerical example and conclusions.

Notation: The $n \times n$ identity matrix is $I_{n \times n}$, and the $n \times m$ matrix with all elements equal to 1 is $1_{n \times m}$. The vectorised form of a matrix $A = [a_1 \ldots a_n]$ is vec$(A) := [a_1^\top \ldots a_n^\top]^\top$ and $A \otimes B$ is the Kronecker product. The Moore–Penrose inverse of a matrix $A$ is denoted as $A^\dagger$. The Euclidean norm is $\|x\|$ and, for a matrix $Q$, $Q \succ 0$ ($Q \succeq 0$) indicates that $Q$ is positive definite (semidefinite) and $\|x\|^2 := x^\top Q x$. The notation $\lim_{\alpha \to 0^+} f(x)$ denotes the one-sided limit of $f(x)$ as $x$ approaches $a$ from the left and $\text{mat}(x, [r, c])$ denotes the columnwise operation to reshape a vector $x$ of appropriate length into an $r$-by-$c$ matrix.

2 Problem description

2.1 System model and feedback information

We assume a system with linear discrete time dynamics:

\begin{align}
    x_{k+1} &= Ax_k + Bu_k + Dw_k, \\
    y_k &= Cx_k + v_k,
\end{align}
\[ z_k = \gamma_k y_k, \quad (1c) \]

where \( x_k \in \mathbb{R}^{n_x}, u_k \in \mathbb{R}^{n_u}, y_k \in \mathbb{R}^{n_y}, z_k \in \mathbb{R}^{n_z}, w_k \in \mathbb{R}^{n_w}, v_k \in \mathbb{R}^{n_v} \) are the state, control input, sensor measurement, measurement information received by the controller, disturbance and the measurement noise, respectively. The disturbance, measurement noise and packet loss sequences, \( \{ w_k \}_{k=0}^{\infty}, \{ v_k \}_{k=0}^{\infty} \) and \( \{ \gamma_k \}_{k=0}^{\infty} \) are assumed to be i.i.d. with
\[
\mathbb{E}\{ w_k \} = 0, \quad \mathbb{E}\{ w_k w_k^\top \} =: \Sigma_w \geq 0, \\
\mathbb{E}\{ v_k \} = 0, \quad \mathbb{E}\{ v_k v_k^\top \} =: \Sigma_v \geq 0, \\
\mathbb{P}\{ \gamma_k = 0 \} = 1 - \lambda, \quad \mathbb{P}\{ \gamma_k = 1 \} =: \lambda.
\]
The variable \( \gamma_k \in \{0, 1\} \) indicates whether sensor data at the \( k \)th sampling instant is received by the controller. The information assumed available to the controller at the \( k \)th time instant consists of \( \{ w_i \}_{i=0}^{k-1}, \{ (z_i, \gamma_i) \}_{i=0}^{k} \), the initial mean \( \mathbb{E}\{ x_0 \} =: \hat{x}_0 \), and model state covariance \( \mathbb{E}\{ (x_0 - \hat{x}_0)(x_0 - \hat{x}_0)^\top \} =: \Sigma_0 \).

We define the information sets
\[
\mathcal{I}_k := \{ \mathcal{I}_{k-1}, \{ z_k, \gamma_k \} \}, \quad \mathcal{U}_k := \{ \mathcal{U}_{k-1}, u_k \},
\]
for all \( k \geq 0 \), where \( \mathcal{I}_{-1} := \{ \hat{x}_0, \Sigma_0 \}, \mathcal{U}_{-1} := \{ \} \). Finally, we define conditional expectation operators as
\[
\mathbb{E}_k \{ \} := \mathbb{E}\{ \cdot | \mathcal{U}_{k-1}, \mathcal{I}_{k-1} \}, \quad \mathbb{E}_0 \{ \} := \mathbb{E}\{ \cdot \}.
\]

**Assumption 1** The pair \((A, B)\) is stabilisable, and \((A, C)\) is detectable.

### 2.2 Optimal control problem

We consider the problem of minimising the discounted sum of expected future values of \( \| x_k \|^2_Q + \| u_k \|^2_R \), where \( Q \geq 0 \) and \( R > 0 \). The minimisation is performed subject to a constraint on the discounted sum of second moments of an auxiliary output, defined for a given matrix \( H \) by \( \xi_k = H x_k \), so the optimal control problem is
\[
\min_{\theta} \sum_{k=0}^{\infty} \beta^k \mathbb{E}\{ \| x_k \|^2_Q + \| u_k \|^2_R \}
\]
subject to
\[
\sum_{k=0}^{\infty} \beta^k \mathbb{E}\{ \| H x_k \|^2 \} \leq \epsilon. \tag{2}
\]

Here \( \beta_1 \in (0, 1) \) and \( \beta_2 \in (0, 1) \) are discount factors and \( \epsilon \) is a given constraint threshold. For simplicity (with the understanding that the general case of \( \beta_1 \neq \beta_2 \) can be treated similarly), we set \( \beta_1 = \beta_2 = \beta \) for the remainder of the paper with the exception of Sections 7.2 and 9.

**Remark 1** The constraint in (2) can be interpreted as an energy constraint, for example, on signals transmitted over a network, and we note that constraints involving first order moments can be handled with minor modifications to our approach. The discount factor in (2) weights performance closer to the initial time more heavily than future behaviour. In contrast to average constraints over an infinite horizon (e.g. [Leong and Dey, 2017]), this discounted constraint allows for a trade-off between short-term and long-term behaviours and indicates preferential allocation of the energy budget towards time instants nearer the initial time.

We will employ a receding horizon control law based on a finite-horizon control policy in the form
\[
\begin{align*}
\theta_{i,k} & = \kappa_i(\theta_k, \mathcal{U}_{k+i-1}, \mathcal{I}_{k+i}) \quad (3)
\end{align*}
\]
where \( u_{i,k} \) for \( i = 0, 1, \ldots \) is the prediction of \( u_{k+i} \) at time instant \( k \), and \( \theta_k \) is a vector of controller parameters at time instant \( k \). The dependence of \( \kappa_i(\cdot) \) on the sets \( \mathcal{U}_{k+i-1} \) and \( \mathcal{I}_{k+i} \) ensures causality and the dependence on \( \theta_k \) is chosen so that the optimal parameter vector, denoted \( \theta^*_k \), will be the solution of a convex problem.

**Assumption 2** (i) The probability, \( \lambda \), of successfully receiving sensor measurements is known and time-invariant. (ii) When \( \theta^*_k \) is computed, \( (z_{k+i}, \gamma_{k+i}) \) are unknown for all \( i \geq 0 \).

Thus \( \theta^*_k \) is computed online prior to the \( k \)th sampling instant using knowledge of \( \mathcal{U}_{k-1} \) and \( \mathcal{I}_{k-1} \), while \((z_k, \gamma_k)\) is known when the control law is implemented, via
\[
\begin{align*}
\mathbf{u}_k & = \mu_0(\theta^*_k, \mathcal{U}_{k-1}, \mathcal{I}_k).
\end{align*}
\]

In order to address (2), we define an MPC optimisation problem to be solved at time instant \( k \) as follows
\[
\begin{align*}
\theta^*_k & = \arg\min_{\theta_k} \sum_{i=0}^{\infty} \beta^i \mathbb{E}_k \{ \| x_{i,k} \|^2_Q + \| u_{i,k} \|^2_R \}
\end{align*}
\]
subject to
\[
\sum_{i=0}^{\infty} \beta^i \mathbb{E}_k \{ \| H x_{i,k} \|^2 \} \leq \mu_k. \tag{4}
\]

Here \( \mu_0 = \epsilon \), and \( \mu_k \) for \( k > 0 \) is chosen as described in Section 6 to ensure that (4) is recursively feasible and that the closed loop system satisfies the constraint in (2).

### 3 Controller design

#### 3.1 Controller parameterisation

Consider the output feedback control law defined by an observer and an affine feedback law:
\[
\begin{align*}
\dot{x}_k = A \hat{x}_{k-1} + Bu_{k-1}, \\
\hat{x}_k = \hat{x}_k + \gamma_k M (y_k - C \hat{x}_k), \\
u_k = K \hat{x}_k + c_k
\end{align*}
\]
where \( \dot{x}_0 = \mathbb{E}\{ x_0 \} \), where \( \hat{x}_k \) and \( \hat{x}_k \) are the prior estimate and the posterior estimate of \( x_{k+1} \), defined as \( \hat{x}_k := \mathbb{E}_k \{ x_k \} \) and \( \hat{x}_k := \mathbb{E}_k \{ x_k \} \). Equations (5a) and (5b) are referred to as a prediction step and an update step respectively. Together with the assumption on \( w_k \) and \( v_k \), these two steps ensure the estimator is unbiased, namely, \( \mathbb{E}\{ x_k \} = \mathbb{E}\{ \hat{x}_k \} = \mathbb{E}\{ \hat{x}_k \} \).

If we fix the gains \( M \) and \( K \) and choose the decision variables in problem (4) as \( \theta_k = \{ \gamma_i k, \ldots, \gamma_{N-1} k \} \) for a fixed \( N \), with the predicted control sequence defined as \( u_{i,k} = u_{i,k} = \ldots = u_{i,N-1,k} = \ldots = u_{i,k} \),
Instead we use a parameterisation that allows optimisation of the dependence of predicted control inputs on future realisations of model uncertainty. This allows the predicted probability distributions of states and control inputs to be controlled explicitly, thus providing flexibility to balance conflicting requirements for performance and constraint satisfaction. Analogously to the case of predicted control laws in which state feedback gains are optimisation variables [Löfberg (2003); Goulart et al. (2006)], the cost and constraints of problem (4) become nonconvex if time-varying gains $M$, $K$ are chosen as decision variables. However, if the predicted control sequence is parameterised in terms of affine functions of the future output measurements received by the controller [Ben-Tal et al. (2009)], then the first and second moments of predicted states and inputs are convex functions of controller parameters. By allowing arbitrary linear dependence of $\kappa_i(\cdot)$ on the received sensor measurements, this approach makes it possible to optimise the future control sequence at every sampling instant, including those at which information from sensors is lost.

Therefore, let the $i$-step ahead predicted control input $u_{i|k}$ be defined for all $i = 0, 1, \ldots$ as

\[
u_{i|k} = K\tilde{x}_{i|k} + c_{i|k} + d_{i|k},
\]

\[
de_{i|k} = \gamma_i L_{i|1;k}(y_{i|k} - C\tilde{x}_{i|k}) + \gamma_{i|k} L_{i|1;k}(y_{i|k} - C\tilde{x}_{i|k})
\]

\[
\vdots + \gamma_{i|k} L_{i|1;k}(y_{i|k} - C\tilde{x}_{i|k}),
\]

\[
\hat{x}_{i+1|k} = A\hat{x}_{i|k} + B\hat{u}_{i|k} + \gamma_i AM(y_{i|k} - C\tilde{x}_{i|k}),
\]

\[
\hat{x}_{i+1|1} = A\hat{x}_{i|1} + B\hat{u}_{i|1} + \gamma_iAM(y_{i|1} - C\tilde{x}_{i|1}),
\]

\[
\Psi_{i|k} := A(I - \gamma_{i|k}MC)\Psi_{i|k} + Dv_{i|k} + Dw_{i|k},
\]

\[
x_{i+1|k} = \Psi_{i|k}(\tilde{x}_{i|k} - \hat{x}_{i|k}) - \gamma_{i|k} AMv_{i|k} + Dw_{i|k},
\]

\[
S_{\Psi} := \prod_{i=0}^{N-1} \Psi_{i|k}, T_{(\Psi, B)} := \prod_{i=0}^{N-1} \Psi_{i|k}B + B 0.
\]

The gains $K$, $M$ in (6a-c) are fixed and determined offline, satisfying the following assumption.

**Assumption 3** Gains $K$ and $M$ are chosen so that $\xi_{i+1} = (A + BK)\xi_i$ is asymptotically stable and $\xi_{i+1} = A(I - \gamma_i MC)\xi_i$ is mean-square stable (MSS) [Kushner, 1974].

**Remark 2** Matrices $K$ and $M$ exist satisfying Assumption 3 if Assumption 1 holds and if the probability, $\lambda$, of successfully receiving a sensor measurement is greater than a critical value (e.g. Sinopoli et al., 2004). Suitable choices for $K$, $M$ are the optimal gains for (4) in the absence of constraints, or the certainty equivalent LQ feedback gain for a problem with state and control weighting matrices $Q$ and $R$ and the steady state Kalman filter gain [Sinopoli et al., 2004]. Note that time-varying gains $K_k$, $M_k$ can be used within the framework of this paper, provided the dependence on $\gamma_k$ is known in advance.

### 3.2 Controller implementation

The control law is implemented as follows:

(i) Given $U_{k-1}$ and $L_{k-1}$, solve problem (4) for $\theta^*_k$;

(ii) Given $\gamma_k$ and $z_k = \gamma_k y_k$:

(a) apply the control input

\[
u_k = K\tilde{x}_k + c_{0|k} + \gamma_k L_{0,0|k}(y_k - C\tilde{x}_k),
\]

(b) update the state estimate

\[
\hat{x}_{k+1} = A\hat{x}_k + B\nu_k + \gamma_k AM(y_k - C\tilde{x}_k).
\]

**Note** that this receding horizon control law includes (5) as a special case, since $u_k$ and $\hat{x}_{k+1}$ in step (ii) are equal to their counterparts in (5) if $(c_{0|k}, L_{0,0|k}) = (c_k, KM)$.

### 4 Predicted state and control sequences

To simplify notation, we define vectorised sequences, with $x_k$ denoting the vectorised true state sequence $\{x_{i|k}\}^{N-1}_{i=0}$, $\tilde{x}_k$ the estimated state sequence $\{\tilde{x}_{i|k}\}^{N-1}_{i=0}$, $u_k$ the predicted control sequence $\{u_{i|k}\}^{N-1}_{i=0}$, $c_k$ the predicted control perturbations $\{c_{i|k}\}^{N-1}_{i=0}$, $w_k$ the disturbance sequence $\{w_{i|k}\}^{N-1}_{i=0}$, $v_k$ the sensor noise sequence $\{v_{i|k}\}^{N-1}_{i=0}$, and $\xi_k$ the future innovation sequence $\{\xi_{i|k}\}^{N-1}_{i=0}$ at time instant $k$. For a given sequence of matrices $\{\Psi_{i|k}\}^{N-1}_{i=0}$ and matrix $B$ let

\[
S_{\Psi} := \prod_{i=0}^{N-1} \Psi_{i|k}, T_{(\Psi, B)} := \prod_{i=0}^{N-1} \Psi_{i|k}B + B 0.
\]
where $\prod_{i=m}^{n} \Psi_{i[k]} := \Psi_{m[k]} \cdots \Psi_{n[k]}$ for $m \geq n$, and define

$$L_k := \begin{bmatrix}
L_{00} & L_{01} & \cdots & L_{10} \\
L_{10} & \ddots & \ddots & \vdots \\
\vdots & \ddots & \ddots & \ddots \\
L_{N1} & \cdots & L_{N0} & L_{00}
\end{bmatrix},$$

$$K_k := \text{diag}(\gamma_0, \ldots, \gamma_{N-1}), I_{n_y \times n_y},$$

$$K := I_{n_y \times n_y} \otimes K, M := I_{n_x \times n_x} \otimes M$$

and

$$\begin{align}
\mathbb{E}_k \{ \pi_k \pi_k^T \} &= [I \mathbf{0}_k] \Omega_k \begin{bmatrix} I \\ \mathbf{0}_k \end{bmatrix}^T, \\
\mathbb{E}_k \{ u_k u_k^T \} &= (K \pi_k + c_k)(K \pi_k + c_k)^T + [0 \mathbf{L} + K \Pi_k] \Omega_k \begin{bmatrix} 0 \\ (L_k + K \Pi_k)^T \end{bmatrix}^T.
\end{align}$$

**Proof:** From (11), we have $\mathbb{E}_k \{ x_k - \hat{x}_k \} = 0$. Therefore $\zeta_k = \Gamma_k C (x_k - \hat{x}_k) + \Gamma_k v_k$ implies $\mathbb{E}_k \{ \zeta_k \} = 0$ and (14a,b) follow from the expectations of (12a,b). To determine the second moments of $x_k$ and $u_k$, let

$$X_k := \mathbb{E}_k \begin{bmatrix} x_k - \hat{x}_k \\ x_k \end{bmatrix} \begin{bmatrix} x_k - \hat{x}_k \\ x_k \end{bmatrix}^T.$$

Then from (11) and (12a) we have

$$X_k = \begin{bmatrix} 0 & 0 \\ \Pi_k \pi_k^T \\ 0 \mathbf{0}_k \end{bmatrix} \Omega_k \begin{bmatrix} 0 & 0 \\ I \mathbf{0}_k \end{bmatrix}^T.$$

(15) and (15a,b) follow from $\mathbb{E}_k \{ x_k x_k^T \} = [I \mathbf{I}] X_k [I \mathbf{I}]^T$ and (12a,b), respectively.

Since $\pi_k, \Pi_k$ are linear in $\theta_k = (c_k, L_k)$ and $\Omega_k$ is independent of $\theta_k$ (as will be shown by (17)), it is clear from (14a,b) and (15a,b) that the first moments of the predicted state and input sequences are linear in $\theta_k$ while their second moments are quadratic functions of $\theta_k$.

To determine $\Omega_k$, note that $x_k - \hat{x}_k$ and $\zeta_k$ can be written

$$x_k - \hat{x}_k = F(\Gamma_k)g_k, \quad \zeta_k = G(\Gamma_k)q_k, \quad q_k := \begin{bmatrix} v_k \\ w_k \end{bmatrix},$$

with $F(\Gamma_k) := [S_\Phi - T(\Phi, \mathcal{A})M_\Gamma k \sigma_{k}]$ and $G(\Gamma_k) := \Gamma_k C F(\Gamma_k) + [0 \mathbf{I} k \mathbf{0}]$. Hence, by the law of total expectation,

$$\Omega_k = \sum_j \begin{bmatrix} F(\Gamma^{(j)}) & G(\Gamma^{(j)}) \end{bmatrix} \mathbb{E}_k \{ q_k q_k^T \} \begin{bmatrix} F(\Gamma^{(j)}) & G(\Gamma^{(j)}) \end{bmatrix}^T \mathbb{P}(\Gamma = \Gamma^{(j)}),$$

(17)

where $\mathbb{E}_k \{ q_k q_k^T \}$ is the block-diagonal matrix:

$$\mathbb{E}_k \{ q_k q_k^T \} = \text{diag}(\Sigma_k, \Sigma_v, \Sigma_u),$$

and where $\Gamma^{(j)}$ for $j = 1, \ldots, 2^N$ enumerates the $2^N$ matrices with binary-valued diagonal elements defined by $\Gamma(1) = \mathbf{I}$, $\Gamma(2) = \mathbf{I}$, $\Gamma(3) = \mathbf{I}$, $\ldots$, $\Gamma(2^N - 1) = \mathbf{I}$.

**Remark 4** The matrix $\Omega_k$ in (17) can be computed conveniently via

$$\text{vec}(\Omega_k) =$$

\[
\left( \sum_j \begin{bmatrix} F(\Gamma^{(j)}) \\ G(\Gamma^{(j)}) \end{bmatrix} \otimes \begin{bmatrix} F(\Gamma^{(j)}) \\ G(\Gamma^{(j)}) \end{bmatrix} \right) \mathbb{P}(\Gamma_k = \Gamma^{(j)}) \text{vec}\left( [\Sigma_k, \Sigma_w] \right)
\]
where the first term on the RHS can be determined offline given the probability distribution of \( \gamma_k \). This allows \( \Omega_k \) to be computed online using the current value of \( \Sigma_k \) with a single matrix-vector multiplication.

Using the same arguments as the proof of Proposition 3, it can be verified that

\[
X_{N|i} := \mathbb{E}_k \left\{ \begin{bmatrix} x_{N|i} - \hat{x}_{N|i} \\ \hat{x}_{N|i} \end{bmatrix} \right\} = \begin{bmatrix} 0 & 0 \\ \pi_{N|i} \pi_{N|i}^T & I \end{bmatrix} + \begin{bmatrix} 0 & 0 \\ 0 & \Pi_{N|i} \end{bmatrix} \Omega_{N|i} \begin{bmatrix} 0 & 0 \\ \Pi_{N|i} \end{bmatrix}^T
\]

where \( \Omega_{N|i} := \sum_j \begin{bmatrix} F_N(\Gamma^{(j)}) \\ G(\Gamma^{(j)}) \end{bmatrix} \otimes \begin{bmatrix} \Sigma_k \\ \Sigma_w \end{bmatrix} \mathbb{P}(\Gamma_k = \Gamma^{(j)}) \) with \( F_N(\Gamma_k) := [S_N - T_N(\Phi_\gamma) M] G(\Gamma^{(j)}) \) and \( S_N, T_N(\Phi_\gamma), T_N(\Phi_B) \) being defined (analogously to \( S_N \) and \( T_N(\Phi_A), T_N(\Phi_B) \) in terms of \( \Phi, A \) and \( B \).

5 MPC optimisation

In this section, we formulate the MPC optimisation problem to be repeatedly solved online, using the expressions of the first and second moments of the predicted state and control sequences derived in Section 4.

First note that the objective in (4) can be written

\[
\sum_{i=0}^{\infty} \beta^i \mathbb{E}_k \left\{ \| x_{i|k} \|^2 + \| u_{i|k} \|^2 \right\} = \text{tr}(Q\beta X_k) + \text{tr}(R\beta U_k) + f_N(\theta_k, \hat{x}_k, \Sigma_k) \tag{18}
\]

where \( X_k \) is given by (16), and

\[
Q_\beta := 1_{2\times2} \otimes \text{diag} \{ Q, \beta Q, \ldots, \beta^{N-1} Q \}, \\
R_\beta := \text{diag} \{ R, \beta R, \ldots, \beta^{N-1} R \}, \\
U_k := \mathbb{E}_k \{ u_k u_k^T \},
\]

\[
f_N(\theta_k, \hat{x}_k, \Sigma_k) := \sum_{i=0}^{\infty} \beta^i \mathbb{E}_k \left\{ \| x_{i|k} \|^2 + \| u_{i|k} \|^2 \right\}.
\]

Since \( Q_\beta \succeq 0 \) and \( R_\beta > 0 \), the term \( \text{tr}(Q_\beta X_k) + \text{tr}(R_\beta U_k) \) in (18) can be expressed as a convex quadratic function of \( \theta_k = (c_k, L_k) \) using (15b) and (16). To determine the terminal term, \( f_N(\theta_k, \hat{x}_k, \Sigma_k) \), let

\[
P_k := \sum_{i=0}^{\infty} \beta^i X_{i|k}, \tag{19}
\]

where

\[
X_{i|k} := \mathbb{E}_k \left\{ \begin{bmatrix} x_{i|k} - \hat{x}_{i|k} \\ \hat{x}_{i|k} \end{bmatrix} \right\}.
\]

Then for \( i \geq N \) we have

\[
X_{i+1|k} = \mathbb{E} \{ \tilde{\Psi}(\gamma) X_{i|k} \tilde{\Psi}(\gamma)^T \} + \mathbb{E} \{ \tilde{D}(\gamma) \Sigma \Sigma_w \tilde{D}(\gamma)^T \}
\]

(20)

where

\[
\tilde{\Psi}(\gamma) := \begin{bmatrix} A(I - \gamma MC) & 0 \\ \gamma AMC & \Phi \end{bmatrix}, \quad \tilde{D}(\gamma) := \begin{bmatrix} -\gamma AM & D \\ \gamma AM & 0 \end{bmatrix},
\]

and \( \gamma \) is essentially \( \gamma_{ijk} \) with subscripts being omitted for simplicity. Hence

\[
\mathbb{E} \{ \tilde{\Psi}(\gamma) P_k \tilde{\Psi}(\gamma)^T \} = \sum_{i=N}^{\infty} \beta^i (X_{i+1|k} - \mathbb{E} \{ \tilde{D}(\gamma) \Sigma \Sigma_w \tilde{D}(\gamma)^T \})
\]

\[
= \beta^{-1} (P_k - \beta^N X_{N|i}) - \frac{\beta^N}{1 - \beta} \mathbb{E} \{ \tilde{D}(\gamma) \Sigma \Sigma_w \tilde{D}(\gamma)^T \}
\]

where the terminal term \( f_N(\theta_k, \hat{x}_k, \Sigma_k) \) in (18) is equal to \( \text{tr}(Z_1 P_k) \) where \( Z_1 := \begin{bmatrix} Q & Q + K^T R_K \end{bmatrix} \) and \( P_k \) is the solution to the stochastic Lyapunov equation

\[
P_k = \beta \mathbb{E} \{ \tilde{\Psi}(\gamma) P_k \tilde{\Psi}(\gamma)^T \} + \beta^N X_{N|i} + \frac{\beta^{N+1}}{1 - \beta} \mathbb{E} \{ \tilde{D}(\gamma) \Sigma \Sigma_w \tilde{D}(\gamma)^T \}. \tag{21}
\]

Lemma 5 By Assumption 3, the linear system \( \xi_{i+1} = \Psi(\xi) \xi_i \) is mean-square stable.

Proof: In Assumption 3, it is assumed that \( \Psi(\gamma) := A(I - \gamma MC) \) is MSS and \( \Phi \) is asymptotically stable, which are equivalent to

\[
\exists \exists \xi_1 = \xi_1^T \succ 0, \xi_2 = \begin{bmatrix} \xi_1 & \xi_2 \end{bmatrix} \Psi(\gamma) \xi_1^T \xi_2 > 0,
\]

\[
\exists \xi_2 = \xi_2^T \succ 0, \xi_2 = \Phi \xi_2 \Phi^T \succ 0,
\]

respectively. It can be shown that \( \begin{bmatrix} \xi_1 \\ \xi_2 \end{bmatrix} \) satisfies

\[
\begin{bmatrix} \xi_1 \\ \xi_2 \end{bmatrix} = \begin{bmatrix} \xi_1 \\ \xi_2 \end{bmatrix} \Psi(\gamma) \xi_1^T \xi_2 > 0,
\]

and this implies Lemma 5.

Let \( \bar{X} \geq 0 \) denote the steady state solution to (20). Then, by Lemma 5, it is ensured that \( X_{i|k} \) converges to \( \bar{X} \) as \( i \to \infty \) for any \( X_{N|i} \geq 0 \) and \( k \geq 0 \). Also, Lemma 5 implies \( \beta^i \tilde{\Psi}(\gamma) \) is MSS since \( \beta \in (0, 1) \) and thus \( P_k \) in (19) is well defined and finite.

Re-writing the constraints of problem (4) using the matrix, \( X_{i|k} \), of second moments yields the condition

\[
\sum_{i=0}^{\infty} \beta^i \text{tr}(1_{2\times2} \otimes H^T H) X_{i|k} \leq \mu_k,
\]

which is equivalent to the constraint

\[
\text{tr}(H_\beta X_k) + \text{tr}(Z_2 P_k) \leq \mu_k, \tag{22}
\]

where \( H_\beta := 1_{2\times2} \otimes \text{diag} \{ H^T H, \beta H^T H, \ldots, \beta^{N-1} H^T H \} \).
and \( Z_2 := 1_{2 \times 2} \otimes H^\top H \).

The expressions for the cost and constraints in (18), (21), (22) allow the optimisation (4) defining \( \theta_k^* \) to be formulated as

\[
(\theta_k^* , P_k^* ) := \arg \min_{\theta_k , P_k} \text{tr}(Q_\beta X_k) + \text{tr}(R_\beta U_k) + \text{tr}(Z_1 P_k)
\]

s.t. (22),

\[
P_k \geq \beta E \{ \tilde{\Psi}(\gamma) P_k \tilde{\Psi}^\top(\gamma) \} + \beta^N X_{N\mid k}
\]

\[
+ \frac{\beta^{N+1}}{1-\beta} E \{ \tilde{D}(\gamma) \left[ \Sigma_\nu \Sigma_\omega \right] \tilde{D}^\top(\gamma) \}.
\]

(23)

**Remark 6** The relaxation of (21) as a linear matrix inequality (LMI) in problem (23) does not introduce any conservativeness as it can be shown that there always exists a solution that satisfies this LMI with equality. Suppose there does not exist such a solution to problem (23). Let \((\theta_k^* , P_k^0)\) be a minimiser and

\[
P_k^{(1)} := \beta E \{ \tilde{\Psi}(\gamma) P_k^{(0)} \tilde{\Psi}^\top(\gamma) \} + \beta^N X_{N\mid k}
\]

\[
+ \frac{\beta^{N+1}}{1-\beta} E \{ \tilde{D}(\gamma) \left[ \Sigma_\nu \Sigma_\omega \right] \tilde{D}^\top(\gamma) \}.
\]

Then \( P_k^{(0)} \succ P_k^{(1)} \) and \((\theta_k^* , P_k^{(1)})\) also satisfies the LMI with strict inequality because of the supposition. Continuing this procedure, we will have that \( \forall i \geq 1 P_k^{(i)} \succ P_k^{(i+1)} \) and that \((\theta_k^* , P_k^{(i+1)})\) is a minimiser and is at least as good as \((\theta_k^* , P_k^{(i)})\) in terms of values of the cost function. From the mean-square stability of \( \beta^+ \tilde{\Psi}(\gamma) \), it follows that \( P_k^* = \lim_{i \to \infty} P_k^{(i)} \) exists and satisfies this LMI with equality. This contradicts the supposition and proves our argument. Therefore, we let \((\theta_k^* , P_k^*)\) be the minimiser that satisfies this LMI constraint with equality. Note that if \( \left[ \begin{array}{cc} Q_k^+ & 0 \\ 0 & K^+ \end{array} \right] \) has full column rank, \((\theta_k^* , P_k^*)\) is unique.

6 Closed loop properties

This section considers the performance of the closed loop system (1) with the control law of Section 3.2. We use the solution \( \theta_k^* = \{ c_k^* , L_k^* \} \) of (4) at time instant \( k \) to construct a feasible, but possibly suboptimal, solution for (4) at time instant \( k+1 \) (i.e. given \( U_k , I_k \)), which we denote \( \theta_{k+1}^* := \{ c_{k+1}^* , L_{k+1}^* \} \), where

\[
c_{k+1}^* := T c_k^* + \bar{L}_0 \gamma L_k(y_k - C x_k), \quad \text{(24a)}
\]

\[
L_{k+1}^* := \left[ \begin{array}{cccc} L_{1\mid k}^* & \cdots & \cdots & \cdots \\ \vdots & \ddots & \vdots & \vdots \\ L_{N-1\mid k}^* & \cdots & L_{N-1,N-1\mid k}^* & 0 \\ 0 & \cdots & 0 & 0 \end{array} \right]. \quad \text{(24b)}
\]

In (24a), \( T \) is a matrix such that

\[
T \left[ \begin{array}{c} c_0 \\ c_1 \\ \vdots \\ c_{N-1} \end{array} \right] = \left[ \begin{array}{c} c_1 \\ \vdots \\ c_{N-1} \end{array} \right], \quad \text{and} \quad \bar{L}_0 := \left[ \begin{array}{c} L_{1\mid 0\mid k}^* \\ \vdots \\ L_{N-1,0\mid k}^* \end{array} \right].
\]

Following Yan et al. (2018), we define the constraint threshold \( \mu_k \) in (4) \( \forall k > 0 \) in terms of \( \theta_k^* \) as

\[
\mu_k := \left\{ \begin{array}{ll} \epsilon , & k = 0 \\ \text{tr}(H_\beta X_k) + \text{tr}(Z_2 P_k^o) , & k > 0 \end{array} \right. \quad \text{(25)}
\]

where

\[
P_k^o := \beta E \{ \tilde{\Psi}(\gamma) P_k^o \tilde{\Psi}^\top(\gamma) \}
\]

\[
+ \beta^N X_{N\mid k}^o + \frac{\beta^{N+1}}{1-\beta} E \{ \tilde{D}(\gamma) \left[ \Sigma_\nu \Sigma_\omega \right] \tilde{D}^\top(\gamma) \}.
\]

(26a)

\[
X_k^o := \left[ \begin{array}{cc} 0 & 0 \\ 0 & \pi_k \pi_k^\top \end{array} \right] + \left[ \begin{array}{cc} I \otimes & I \otimes \\ 0 \otimes \pi_k \otimes & 0 \otimes \pi_k \otimes \end{array} \right] \Omega_k
\]

\[
\times \left[ \begin{array}{cc} I \otimes & I \otimes \\ 0 \otimes \pi_k \otimes & 0 \otimes \pi_k \otimes \end{array} \right], \quad \text{(26b)}
\]

\[
X_{N\mid k}^o := \left[ \begin{array}{cc} 0 & 0 \\ 0 & \pi^o \otimes \pi^o \end{array} \right] + \left[ \begin{array}{cc} I \otimes & I \otimes \\ 0 \otimes \pi^o \otimes & 0 \otimes \pi^o \otimes \end{array} \right] \Omega_{N\mid k}
\]

\[
\times \left[ \begin{array}{cc} I \otimes & I \otimes \\ 0 \otimes \pi^o \otimes & 0 \otimes \pi^o \otimes \end{array} \right]. \quad \text{(26c)}
\]

with \( \pi_k^o := S \Phi \tilde{x}_k + T(\Phi_B) c_k^o , \ \pi_{N\mid k}^o := S \Phi \tilde{x}_k + T(\Phi_B) c_k^o , \ \pi_{N\mid k}^o := T(\Phi_B) L_k^o + T(\Phi_A) M \)

\[
\pi_{N\mid k}^o := S \Phi \tilde{x}_k + T(\Phi_B) c_k^o , \ \pi_{N\mid k}^o := T(\Phi_B) L_k^o + T(\Phi_A) M.
\]

Combining (22) and (25), we can see that the design of \( \mu_k \) enforces feasibility of \( \theta_k^* \) \( \forall k > 0 \) and therefore ensures recursive feasibility of the MPC optimisation without requiring bounds on the noise \( v_k \) and disturbance \( w_k \).

**Theorem 7** If problem (23) is feasible at \( k = 0 \), then (23) remains feasible for all \( k > 0 \) and the state of (1) under the control law of Section 3.2 satisfies

\[
\sum_{k=0}^{\infty} \beta^k E \{ \| H x_k \|^2 \} \leq \epsilon. \quad \text{(27)}
\]

**Proof:** The definition (25) of \( \mu_k \) trivially ensures feasibility for all \( k > 0 \). The definitions (24a,b) ensure that, at time instant \( k \) (given \( U_{k-1} , I_{k-1} \)), the distributions of the state and control sequences \( \{ x_{i+1\mid k} \}_{i=0}^{\infty} \) and \( \{ u_{i+1\mid k} \}_{i=0}^{\infty} \) are identical to the distributions of \( \{ x_{i+1\mid k} \}_{i=0}^{\infty} \) and \( \{ u_{i+1\mid k} \}_{i=0}^{\infty} \). Therefore

\[
\sum_{i=0}^{\infty} \beta^i E_k \{ \| H x_{i\mid k} \|^2 \} = \text{tr}(H_\beta X_k) + \text{tr}(Z_2 P_k)
\]

implies

\[
\beta^i E_k \{ \| H x_{i\mid k} \|^2 \} \leq \mu_k - E_k \{ \| H x_{i\mid k} \|^2 \} = \mu_k - E_k \{ \| H x_{i\mid k} \|^2 \}.
\]

Hence the trajectories of the closed loop system satisfy

\[
\sum_{i=0}^{\infty} \beta^i E_k \{ \| H x_{i\mid k} \|^2 \} \leq \mu_k - \lim_{i \to \infty} \beta^i E \{ \mu_{k+i} \} \leq \mu_k
\]

for all \( k \geq 0 \).
Corollary 8 Let \( J_k := J(\theta_k^*, \hat{x}_k, \Sigma_k) \) denote the optimal cost of problem (23). Then, under the control law of Section 3.2, the trajectories of (1) satisfy
\[
\sum_{k=0}^{\infty} \beta^k E\left\{ \|x_k\|_Q^2 + \|u_k\|_R^2 \right\} \leq J_0.
\] (28)

Proof: Using the same argument in the proof of Theorem 7 on the definition of the objective in (23) yields
\[
\beta E_k [J(\theta^*_{k+1}, \hat{x}_{k+1}, \Sigma_{k+1})] = J_k - E_k \left\{ \|x_k\|_Q^2 + \|u_k\|_R^2 \right\},
\] (29)
and since \( J_k \leq J(\theta_k^*, \hat{x}_k, \Sigma_k) \) \( \forall k \) by optimality, the bound in (28) follows. \( \square \)

7 Strengthened stability conditions

In Corollary 8, we provide an upper bound on the discounted closed loop cost. However, this may not prevent closed loop states from gradually diverging if the closed loop cost is heavily discounted. Therefore, in the following, we provide stability conditions in terms of bounds on the undiscouted closed loop cost.

7.1 Closed loop cost bound via a numerical check

In this section, we provide a bound on the averaged undiscounted closed loop cost accumulated over an infinite horizon by running a numerical check on \( \beta \). Specifically, if some condition holds for values of \( \beta \) in \((0, 1)\), we can provide a bound on this closed loop cost that is parametric in \( \beta \). To this end, we first establish an upper bound on \( J_k \).

An upper bound on \( J_k \) can be provided by finding a feasible solution to problem (23), and such a solution can be easily obtained if we eliminate variable \( P_k \) and rewrite problem (23) as follows. Using standard matrix vectorisation identities, (21) can be rewritten as
\[
\text{vec}(P_k) = W_1 \text{vec}(\beta^N X_{N|k} + W_2),
\] (30)
where
\[
W_1 := \left[I - \beta(1 - \lambda) \hat{\Psi}(0) \otimes \hat{\Psi}(0) - \beta \lambda \Theta(1) \otimes \hat{\Psi}(1)\right]^{-1},
\]
\[
W_2 := \frac{\beta^{N+1}}{1 - \beta} \mathbb{E}\left\{ \hat{D}(\gamma) \right\} \left[\begin{array}{c} \Sigma_u \\ \Sigma_w \end{array}\right] \hat{D}^\top(\gamma).
\]
In the definition of \( W_1 \), the matrix inverse is guaranteed to exist since \( \beta^N \hat{\Psi}(\gamma) \) is MSS. From (30) we have
\[
\text{tr}(Z_1 P_k) = \text{vec}(Z_1^\top) \text{vec}(P_k)
\]
\[
= \text{vec}(Z_1^\top) W_1 \text{vec}(\beta^N X_{N|k} + W_2)
\]
\[
= \text{tr}(\beta^N \hat{Z}_1 X_{N|k} + \text{tr}(\hat{Z}_1 W_2)),
\]
where \( \hat{Z}_1 := \text{mat}\left(\text{vec}(Z_1)^\top W_1\right), [2n_x, 2n_x] \). Similarly, \( \text{tr}(Z_2 P_k) = \text{tr}(\beta^N \hat{Z}_2 X_{N|k} + \text{tr}(\hat{Z}_2 W_2)) \), where \( \hat{Z}_2 := \text{mat}\left(\text{vec}(Z_2)^\top W_1\right), [2n_x, 2n_x] \).

Therefore, an equivalent form of problem (23) is given by
\[
\theta_k^* = \arg\min_{\theta_k} \text{tr}(Q \beta X_k) + \text{tr}(R \beta U_k) + \beta^N \text{tr}(\hat{Z}_1 X_{N|k}) + \text{tr}(\hat{Z}_1 W_2)\]
\[\quad + \text{tr}(\hat{Z}_2 W_2),
\]
s.t. \( \text{tr}(H_j X_k) + \beta^N \text{tr}(\hat{Z}_2 X_{N|k}) + \text{tr}(\hat{Z}_2 W_2) \leq \mu_k, \) (31) which has a single convex quadratic constraint.

Minimising the constraint function in (31) over \( \theta_k = \{c_k, L_k\} \) necessarily yields a feasible, but possibly sub-optimal, solution \( \theta_k^* := \{c_k^*, L_k^*\} \). This solution has an explicit form, in which \( c_k^* := K \beta \hat{x}_k \) with
\[
\begin{aligned}
K_c := & \left[ T^{(\Phi, B)}_k H_b T^{(\Phi, B)}_k + \beta^N (T^{(N)}_k)^\top Z_{22,22} T^{(N)}_k \right] \\
& \cdot \left[ T^{(\Phi, B)}_k H_b S_{\Phi, B} + \frac{\beta^N}{2} (T^{(N)}_k)^\top (Z_{22,22} + Z_{22,22}) S_{\Phi, B}^2 \right],
\end{aligned}
\]
\[H_b := \text{diag}\{H^\top H, \beta H^\top H, \ldots, \beta^{N-1} H^\top H\}.\]
The matrix \( Z_{22,22} \) is the bottom-right block of \( Z_2 \) partitioned according to \( \hat{Z}_2 = \left[ \hat{Z}_{21,1}, \hat{Z}_{22,1}, \hat{Z}_{22,2} \right]. \) The expression for \( L_k^* \) is omitted here for simplicity but we note that it is a function of \( \Sigma_k \). Therefore, given \( \theta_k^* = \{c_k^*, L_k^*\} \) we can express the cost function as
\[
J(\theta_k^*, \hat{x}_k, \Sigma_k) = \hat{x}_k^\top P_k \hat{x}_k + \text{tr}(Q \Sigma_k) + s(\beta, \Sigma_k) + \eta.
\] (32)
In (32), \( s(\beta, \Sigma_k) \) is a nonnegative scalar function of \( \beta \) and \( \Sigma_k \), consisting of some terms in \( J(\theta_k^*, \hat{x}_k, \Sigma_k) \) that depend on \( \Sigma_k \) and taking finite values when \( \Sigma_k \) is finite, and \( \eta \) is a constant depending on \( \beta \), \( \Sigma_w \), and other system parameters. The matrix \( P_k \) is given by
\[
P_k := \left( S_{\Phi} + T^{(\Phi, B)}_k K_c \right) Q_b \left( S_{\Phi} + T^{(\Phi, B)}_k K_c \right)^\top + \left( K \left( S_{\Phi} + T^{(\Phi, B)}_k K_c \right) + K_c \right)^\top R \left( K \left( S_{\Phi} + T^{(\Phi, B)}_k K_c \right) + K_c \right) + \frac{\beta^N}{2} \left( S_{\Phi} + T^{(N)}_k K_c \right) \cdot \left( Z_{22,22}^\top + Z_{22,22} \right) \left( S_{\Phi} + T^{(N)}_k K_c \right)^\top
\]
where \( Q_b := \text{diag}\{Q, \beta Q, \ldots, \beta^{N-1} Q\} \) and \( Z_{22,22} \) is the bottom-right block of \( Z_1 = \left[ \hat{Z}_{21,1}, \hat{Z}_{21,2}, \hat{Z}_{22,2} \right]. \)

Moreover, if \( \sigma \) is any scalar such that
\[
P_2 \preceq \sigma Q,
\] (33)
then we have
\[
J_k \leq J(\theta_k^*, \hat{x}_k, \Sigma_k)
\]
\[
\leq \sigma \text{tr}(Q \hat{x}_k \hat{x}_k^\top) + \text{tr}(Q \Sigma_k) + s(\beta, \Sigma_k) + \eta
\]
\[
\leq \sigma \mathbb{E}_k \|x_k\|_Q^2 + s(\beta, \Sigma_k) + \eta,
\] (34)
where the first inequality holds by optimality. From the definitions of \( P_2 \), \( S_{\Phi} \) and \( Q_b \), we have \( P_2 \succeq Q \). This implies that, if a scalar \( \sigma \) exists such that (33) is satisfied, then it must be greater than or equal to 1. Thus the third inequality in (34) follows. A possible choice for \( \sigma \) is the largest eigenvalue of \( P_2 Q^{-1} \) if \( Q > 0 \).

We are now able to identify a parametric bound on the averaged undiscounted closed loop cost accumulated over an infinite horizon in Theorem 9.

Theorem 9 Provided that optimisation (23) is feasible at time instant \( k = 0 \) and a scalar \( \sigma \) exists such that (33) is satisfied, if the discount factor \( \beta \in (0, 1) \) in optimisa-
tion (23) satisfies
\[ \sigma < \frac{1}{1 - \beta} \]  
and the random sequence \{\Sigma_k\}_{k=0}^\infty is upper bounded by some matrix \( \bar{\Sigma} \) (which may depend on \( \Sigma_0 \)), then
\[ \lim_{T \to \infty} \frac{1}{T} \sum_{k=0}^{T-1} \mathbb{E}\left\{ \|x_k\|_Q^2 + \|u_k\|_R^2 \right\} \leq \sup_{\sigma \in \Sigma} \left( s(\beta, \Sigma) + \eta \right) \frac{1}{1 - \beta} - \sigma. \]  
(36)

Proof: From (29), it follows that
\[ \beta E_k \{J_{k+1}\} \leq J_k - E_k \{\|x_k\|_Q^2 + \|u_k\|_R^2\}. \]  
Multiplying both sides of this inequality by \( \beta^{-1} \) and then subtracting \( J_k \) from both sides, we obtain
\[ E_k \{J_{k+1}\} - J_k \leq (\frac{1}{\beta} - 1) J_k - \frac{1}{\beta} E_k \{\|x_k\|_Q^2 + \|u_k\|_R^2\}. \]  
(37)

From (34) and (37), it can be concluded that
\[ \mathbb{E}\left\{ \{J_{k+1}\} - J_k \leq ((\beta^{-1} - 1)\sigma - \beta^{-1}) \mathbb{E}_k \{\|x_k\|_Q^2\} \right\} \]  
\[ - \beta^{-1} \mathbb{E}_k \{\|u_k\|_R^2\} + (\beta^{-1} - 1)(s(\beta, \Sigma_k) + \eta). \]  
(38)

Furthermore, (35) implies that \( \beta^{-1} - (\beta^{-1} - 1)\sigma > 0 \).

Summing both sides of (38) over \( k = 1, \ldots, T - 1 \), dividing by \( T \) and considering the limit as \( T \to \infty \), we get
\[ \lim_{T \to \infty} \frac{1}{T} \sum_{k=0}^{T-1} \mathbb{E}\left\{ \|x_k\|_Q^2 + \|u_k\|_R^2 \right\} \leq \frac{1}{1 - \beta} \sigma - \sigma. \]  
(39)

Also, from \( \Sigma_k \leq \bar{\Sigma} \) \( \forall k \geq 0 \) it follows that
\[ \lim_{T \to \infty} \frac{1}{T} \sum_{k=0}^{T-1} \mathbb{E}\left\{ s(\beta, \Sigma_k) \right\} \leq \sup_{\sigma \in \Sigma} s(\beta, \Sigma). \]  
By the definition of function \( s(\beta, \Sigma) \), this supremum is finite, which completes the proof. \( \square \)

Remark 10 Theorem 9 provides a sufficient condition for (36), which requires \( \Sigma_k \leq \bar{\Sigma} \forall k \geq 0 \). Although not required by mean-square stability alone, this requirement is reasonable in real world applications. It can be ensured, for example, if a common quadratic Lyapunov equation (Lin and Antsaklis [2002]) exists for subsystems of \( r_{i+1} = A(\bar{T} - \gamma_i \text{MC}) \hat{e}_i \), that is, if there exists a matrix \( P = P^\top > 0 \) such that
\[ P - A^\top PA > 0, \]  
\[ P - (A - \text{MC})^\top P(A - \text{MC}) > 0 \]  
(39)
hold simultaneously.

Remark 11 We can relax the condition on \( \Sigma_k \) in Theorem 9 to mean-square stability of \( \Psi(\gamma_k) \) (which is assumed in Assumption 3), if there exist some positive constant \( \alpha \) and symmetric positive definite matrix \( P \) such that
\[ s(\beta, \Sigma_k) \leq \alpha \text{tr}(\Sigma_k P) \]  
\( \forall k \geq 0. \)  
(40)

Condition (40) holds, for example, if a restricted version of the predicted control policy (6a) is used, where \( L_{i,j} \) in (6b) is 0 for all \( i = 0, \ldots, N - 1 \), so that \( L_k \) is not a decision variable in online MPC optimisation problems.

7.2 Closed loop cost bound via asymptotic limits

In this section, we allow the discount factors \( \beta_1 \) and \( \beta_2 \) in the objective and the constraint in problem (2) to differ. The corresponding online optimisation problem defining \( \theta_k^* \) then becomes
\[ \min_{\theta_k, P_{\beta_1,k}, P_{\beta_2,k}} \text{tr}(Q_{\beta_1} X_k) + \text{tr}(R_{\beta_2} U_k) + \text{tr}(Z_1 P_{\beta_1,k}) \]  
s.t. \[ \text{tr}(H_{\beta_2} X_k) + \text{tr}(Z_2 P_{\beta_2,k}) \leq \mu_k, \]  
\[ P_{\beta_1,k} \geq \beta_1 \text{tr}\left\{ \Psi(\gamma_k) P_{\beta_1,k} \Psi(\gamma_k)^\top \right\} + \beta_1^N X_N | k \]  
\[ + \frac{1}{1 - \beta_1} \text{tr}\left\{ \bar{D}(\gamma_k) \left[ \Sigma_v \Sigma_u \right] \bar{D}(\gamma_k)^\top \right\}, \]  
\[ P_{\beta_2,k} \geq \beta_2 \text{tr}\left\{ \Psi(\gamma_k) P_{\beta_2,k} \Psi(\gamma_k)^\top \right\} + \beta_2^N X_N | k \]  
\[ + \frac{1}{1 - \beta_2} \text{tr}\left\{ \bar{D}(\gamma_k) \left[ \Sigma_v \Sigma_u \right] \bar{D}(\gamma_k)^\top \right\}, \]  
(41)

where \( Q_{\beta_1}, R_{\beta_2}, H_{\beta_2} \) are constant matrices similarly constructed to \( Q_\beta, R_\beta, H_\beta \). To analyse the averaged undiscounted closed loop cost accumulated over an infinite horizon, we keep the discount factor \( \beta_2 \) fixed, while taking the left-hand limit of the discount factor in the cost at \( \beta_1 = 1^- \) and then solving the optimisation (41). Considering the limit at \( \beta_1 = 1^- \) implies that \( \beta_1 < 1 \) and hence (41) remains solvable.

We next give a bound on the closed loop cost.

Theorem 12 Provided the optimisation problem (41) is feasible at time instant \( k = 0 \), if the discount factor \( \beta_1 \in (0, 1) \) in (41) is arbitrarily close to 1, then the trajectories of (1) under the control law of Section 3.2 satisfy
\[ \lim_{T \to \infty} \frac{1}{T} \sum_{k=0}^{T-1} \mathbb{E}\left\{ \|x_k\|_Q^2 + \|u_k\|_R^2 \right\} \leq \text{tr}(Z_1 X). \]  
(42)

Proof: We denote the minimiser to (41) with \( \beta_1 \) being arbitrarily close to 1 from the left as \( (\hat{\theta}_k, \hat{P}_k^*, \hat{P}_k^*) \), where \( \hat{P}_k^* \) and \( \hat{P}_k^* \) are the terminal matrices in the predicted cost and the discounted sum constraint, respectively, and \( X_k, U_k, X_{i,k} \) as the quantity of \( X_k, U_k, X_{i,k} \) that correspond to the minimiser \( \hat{\theta}_k^* \), respectively. Let \( \hat{P}_k^{(T)} := \sum_{i=N}^{N+T-1} \hat{X}_{i,k}^* \). Then we have
\[ \hat{P}_k^{(T)} = \sum_{i=N}^{N+T-1} (\hat{X}_{i,k}^* - \hat{X}) + T \hat{X}, \]  
(43)
where \( \hat{X} \), the steady state solution to (20), exists and is unique by Lemma 5. Dividing both sides of (43) by \( T \)
and taking the limit at $T = \infty$, we have

$$\lim_{T \to \infty} \frac{1}{T} \hat{P}_k^T \equiv \lim_{T \to \infty} \frac{1}{T} \sum_{i=0}^{T-1} (\hat{X}_{ik}^* - \bar{X}) + \bar{X} = \bar{X} \quad (44)$$

$$= \lim_{T \to \infty} \frac{1}{T} \sum_{i=0}^{T-1} \sum_{j=0}^{T-1} \beta_1^{i+j} \hat{X}_{ik}^* \equiv \lim_{T \to \infty} \frac{1}{T} \bar{P}_k^* \quad (45)$$

The second equality in (44) follows from the observations that $\hat{X}_{ik}^*$ converges to $\bar{X}$ as $\beta_1 \to 1$, and $T \to \infty$. We next show the second equality in (45) holds. Whenever $(\hat{\theta}_k^*, \bar{P}_k^*)$ satisfies its corresponding LMI constraint in problem (41) with equality, $\bar{P}_k^*$ satisfies (19). Also, since we solve problem (41) after taking the limit at $\beta_1 = 1^-$, each term of that infinite sum on the RHS of (19) is evaluated at $\beta_1 = 1^-$. 

Taking the limit at $\beta_1 = 1^-$, we obtain a slightly different version of (29) as

$$\mathbb{E}_k\{J(\hat{\theta}_{k+1}^*, \hat{x}_{k+1}, \Sigma_{k+1})\} \leq J(\hat{\theta}_k^*, \hat{x}_k, \Sigma_k) - \mathbb{E}_k\{\|x_k\|^2_Q + \|u_k\|^2_R\},$$

and summing both sides of this inequality over $k = 0, \ldots, T - 1$ yields that

$$\sum_{k=0}^{T-1} \mathbb{E}_k\{\|x_k\|^2_Q + \|u_k\|^2_R\} \leq J(\hat{\theta}_0^*, \hat{x}_0, \Sigma_0) - \mathbb{E}_k\{J(\hat{\theta}_T^*, \hat{x}_T, \Sigma_T)\}.$$  

(46)

Dividing both sides of (46) by $T$ and taking the limit at $T = \infty$, we have

$$\lim_{T \to \infty} \frac{1}{T} \sum_{k=0}^{T-1} \mathbb{E}_k\{\|x_k\|^2_Q + \|u_k\|^2_R\} \leq J(\hat{\theta}_0^*, \hat{x}_0, \Sigma_0) - \mathbb{E}_k\{J(\hat{\theta}_T^*, \hat{x}_T, \Sigma_T)\} - \mathbb{E}_k\{J(\hat{\theta}_T^*, \hat{x}_T, \Sigma_T)\}.$$  

(47)

Division by $T$ and taking the limit at $T = \infty$, we have

$$\lim_{T \to \infty} \frac{1}{T} \sum_{k=0}^{T-1} \mathbb{E}_k\{\|x_k\|^2_Q + \|u_k\|^2_R\} \leq \lim_{T \to \infty} \frac{1}{T} J(\hat{\theta}_0^*, \hat{x}_0, \Sigma_0) - \mathbb{E}_k\{J(\hat{\theta}_T^*, \hat{x}_T, \Sigma_T)\}.$$  

(48)

Dividing both sides of (47) by $T$ and taking the limit at $T = \infty$, we have

$$\lim_{T \to \infty} \frac{1}{T} \sum_{k=0}^{T-1} \mathbb{E}_k\{\|x_k\|^2_Q + \|u_k\|^2_R\} \leq \lim_{T \to \infty} \frac{1}{T} \mathbb{E}_k\{J(\hat{\theta}_0^*, \hat{x}_0, \Sigma_0) - \mathbb{E}_k\{J(\hat{\theta}_T^*, \hat{x}_T, \Sigma_T)\}\}.$$  

(49)

This section provides an analysis of the robustness of the MPC controller (9) with respect to uncertainties in the arrival probability, $\mathbb{P}\{\gamma_k = 1\}$, of sensor measurements. Supposing that only a nominal value, $\lambda_n$, of the arrival probability of sensor measurements is given, while its actual value, $\lambda_n := \mathbb{P}\{\gamma_k = 1\}$, is time-invariant but unknown and possibly different from $\lambda_n$, we analyse the impact of the error $\Delta \lambda := \lambda_n - \lambda_n$ on closed loop properties (27)-(28), investigating how one can choose an appropriate value for $\mu_0$ (different from that specified in (25)) so as to increase the robustness margin. To facilitate the following analysis, we define an auxiliary Bernoulli process $\{\tilde{\gamma}_k\}$ with $\mathbb{P}\{\tilde{\gamma}_k = 1\} = \lambda_n$ and $\mathbb{P}\{\tilde{\gamma}_k = 0\} = 1 - \lambda_n$, and we write $\tilde{\gamma}_k \sim B(1, \lambda_n)$ where $B(1, \lambda_n)$ denotes a Bernoulli distribution with a success rate $\lambda_n$.

The nominal value of the arrival probability of sensor measurements is used to formulate optimisation (4) defining $\theta_k^*$ as problem (23) and equivalently as problem (31). More specifically, $\lambda_n$ is used in computing $\Omega_k$ in (17) and the terminal matrix $P_k$ in (21). However, the controller implementation and the updates of $\hat{x}_k, \Sigma_k$ and $\mu_k$ depend on realisations of $\gamma_k$. Hence the random variable $\gamma_k$ in (9), (10), (13) and (24a) assumes values in $\{0, 1\}$ according to the probability distribution $B(1, \lambda_n)$.

Assumption 4  
(i) $(A, \Sigma_{\infty})$ is controllable;  
(ii) The nominal arrival probability $\lambda_n$ is greater than the minimum value (e.g. [35]) such that there exists a gain matrix $M_1$, ensuring the system $\xi_{k+1} = (A - \bar{\gamma}_k AM_1)\xi_k$ is MSS;  
(iii) $|\Delta \lambda|$ is small so that the actual arrival probability $\lambda_a := (\Delta \lambda + \lambda_n)$ is sufficiently large such that there exists a gain matrix $M_2$, ensuring the system $\xi_{k+1} = (A - \gamma_k AM_2)\xi_k$ is MSS.

By exploiting Theorem 5 and Lemma 1 in [35] (which hold by Assumptions 1 and 4(i)), we can show that: if there exists a gain matrix $M'$ such that the system $\xi_{k+1} = (A - \bar{\gamma}_k' AM')\xi_k$ is MSS, where $\{\gamma_k'\}_{k=0}^{\infty}$ is i.i.d. with $\gamma_k' \sim B(1, \lambda')$, then, with the same gain matrix, the system $\xi_{k+1} = (A - \gamma_k' AM')\xi_k$ is also MSS for any Bernoulli process $\{\gamma_k''\}_{k=0}^{\infty}$ that is i.i.d. with $\mathbb{P}\{\gamma_k'' = 1\} \geq \lambda'$ and $\mathbb{P}\{\gamma_k'' = 0\} \leq 1 - \lambda'$. Hence, by Assumption 4, there necessarily exists a gain matrix $M$ such that the systems $\xi_{k+1} = (A - \bar{\gamma}_k AM)\xi_k$ and $\xi_{k+1} = (A - \gamma_k AM)\xi_k$ are MSS.

We next define a simplistic parameterisation of the predicted control sequence as

$$u_{i|k} = K\hat{x}_{i|k} + c_{i|k}, \quad i = 0, \ldots, N - 1, \quad (47a)$$

$$\hat{u}_{i|k} = K\hat{x}_{i|k} \quad (47b)$$

Note that (47) is a special case and a restricted version of (6a,b) with $L_{i|k} = 0 \forall i \geq 0$. Nevertheless, using (47) to formulate optimisation (4) defining $\theta_k^*$, we retain the results on the closed loop system in Sections 6 and 7.
The main result of this section is given below.

**Theorem 13** Under Assumption 4, if parameterisation (47) is used to formulate optimisation (4) defining \( \theta_k^* \), then, in the closed loop operation of the MPC controller (9), a small uncertainty \( \Delta \lambda \) in the arrival probability of sensor measurements results in finite changes to the constraint bound (RHS of (27)) and the cost bound (RHS of (28)) that depend linearly on \( \Delta \lambda \).

**Proof:** We begin the proof with recalling the basis upon which bounds (27) and (28) are derived when we have perfect knowledge about the probability distribution of \( \gamma_k \), namely \( \lambda_n = \lambda_\alpha \). We next introduce intermediate variables \( \mu_k \) to reconstruct this basis when \( \Delta \lambda \neq 0 \) and \( \lambda_n \neq \lambda_\alpha \). Then the proof is completed by quantifying the resulting changes in (27) and (28) from \( \Delta \lambda \) as discounted expectations of \( \mu_k - \mu_k^n \) accumulated over the infinite horizon. The proof of Theorem 7 is based on the observation that, at time instant \( k \), the distributions of the state and control sequences \( \{x_{i+1,k+1}\}_{i=0}^\infty \) and \( \{u_{i+1,k}\}_{i=0}^\infty \) are identical to the distributions of \( \{x_{i+1,k+1}\}_{i=0}^\infty \) and \( \{u_{i+1,k}\}_{i=0}^\infty \). However, this property no longer holds if \( \lambda_n \neq \lambda_\alpha \) since \( \gamma_k \sim B(1, \lambda_n) \) is assumed, whereas \( \gamma_k \sim B(1, \lambda_\alpha) \). Given the definition of \( \mu_k \) in (25), we define a function \( h \) such that

\[
\mu_{k+1} = h(\theta_{k+1}^*, \hat{x}_{k+1}, \Sigma_{k+1}).
\]

We then define an auxiliary variable \( \mu_k^n \) as

\[
\mu_k^n := h(\theta_{k+1}^*, \hat{x}_k^n, \Sigma_k^n) \quad \forall k \geq 0
\]

with

\[
\begin{align*}
\hat{x}_{k+1}^n &= \mathbf{c}_{k+1}^n + L^n_{0} \gamma_k (y_k - C \hat{x}_k), \\
L^n_{k+1} &= L^n_{k}.
\end{align*}
\]

and

\[
\begin{align*}
\mathbf{c}_{k+1}^n &= \mathbf{C}_{k+1}^n \gamma_k (y_k - C \hat{x}_k), \\
\Sigma_{k+1} &= (A - \gamma_k AMC) \Sigma_k (A - \gamma_k AMC)^T + \gamma_k AM \Sigma \Sigma^T A^T + D \Sigma D^T.
\end{align*}
\]

Note that equations (48)-(50) are defined similarly to their counterparts (24a), (10) and (13) respectively. By the definition of function \( h \),

\[
\mu_{k+1} = \sum_{i=0}^\infty \beta^i E_k \{\|H x_{i+1,k}\|^2\},
\]

and the predicted state and corresponding control sequences in (51) are defined by \( \theta_{k+1}^* = \{\mathbf{c}_{k+1}^n, L^n_{k+1}\} \) with initial state \( E_{k+1} \{x_{0,k+1}\} = \hat{x}_{k+1} \) and initial estimation error covariance \( E_{k+1} \{x_{0,k+1} - \hat{x}_{k+1}\} = \Sigma_{k+1} \). Since the probability distribution of \( \gamma_k \) is consistent with the imperfect knowledge of the probability distribution of \( \gamma_k \), we recover the equivalence that, at time instant \( k \), the distributions of the state and control sequences \( \{x_{i+1,k+1}\}_{i=0}^\infty \) and \( \{u_{i+1,k}\}_{i=0}^\infty \) that are associated with \( \theta_{k+1}^*, \hat{x}_{k+1}^n, \Sigma_k^n \) are identical to the distributions of \( \{x_{i+1,k+1}\}_{i=0}^\infty \) and \( \{u_{i+1,k}\}_{i=0}^\infty \). Therefore,

\[
\begin{align*}
E_k \{\mu_{k+1}\} &= \sum_{i=0}^\infty \beta^i E_k \{\|H x_{i+1,k}\|^2\},
\end{align*}
\]

and it follows that

\[
\begin{align*}
\beta E_k \{\mu_{k+1}\} &= \beta E_k \{\mu_{k+1} - \mu_k^n + \mu_k^n\} = \sum_{i=0}^\infty \beta^i E_k \{\|H x_{i+1,k}\|^2\} - E_k \{\|H x_{0,k}\|^2\}.
\end{align*}
\]

Thus, the resulting change in bound (27) from \( \Delta \lambda \) is given by

\[
\sum_{i=0}^\infty \beta^i E_k \{\|H x_{i+1,k}\|^2\} - E_k \{\|H x_{0,k}\|^2\}.
\]

Summing both sides of (52) for \( k = 0, 1, \ldots \), we have

\[
\sum_{k=0}^\infty \beta^k E_k \{\|H x_{k}\|^2\} \leq \mu_0 + \sum_{k=0}^\infty \beta^k \Delta \lambda \Sigma \{\|H x_{k}\|^2\}.
\]

To quantify \( \sum_{k=0}^\infty \beta^k \Delta \lambda \Sigma \{\|H x_{k}\|^2\} \), we next compute \( E_k \{\mu_{k+1} - \mu_k^n\} \). To streamline the presentation of this proof, the detailed algebraic calculation is omitted while noting that \( E_k \{\mu_{k+1}\} \) can be computed using \( X_{N+1}^n \) and \( X_{N+1}^n \) defined in (26) and \( E_k \{\mu_k^n\} \) can be computed using \( X_{N+1}^0 \) and \( X_{N+1}^0 \) that depend on \( \theta_{k+1}^*, \hat{x}_k^n, \Sigma_k^n \). We obtain \( E_k \{\mu_{k+1} - \mu_k^n\} \) as

\[
\begin{align*}
E_k \{\mu_{k+1}\} - E_k \{\mu_k^n\} &= \Delta \lambda \Sigma \{\|H x_{k}\|^2\},
\end{align*}
\]

where

\[
\begin{align*}
\Delta \lambda \Sigma &= \{E_k \{\|H x_{k}\|^2\} - E_k \{\|H x_{0,k}\|^2\}\}.
\end{align*}
\]

It is clear if \( \lambda_n = \lambda_\alpha \), (53) is identical to (27). If parameterisation (47) is used to formulate optimisation (4) defining \( \theta_k^* \), then \( L_{0}^n \gamma_k \), \( \Pi_{k+1}^n \), and \( \Pi_{N+1}^n \) are constant matrices and \( \Sigma_{k+1} \) becomes a linear time-invariant function of \( \Sigma_k \). Moreover, Assumption 4 implies that there exists some matrix \( \Sigma \) such that \( E \{\Sigma\} \leq \Sigma \forall k \geq 0 \). Therefore, the resulting change in bound (27) from \( \Delta \lambda \) can be upper bounded by some finite constant depending linearly on \( \Delta \lambda \) and \( \Sigma \). The analysis of the cost function to show that \( \Delta \lambda \neq 0 \) causes a finite change in the bound (28) is similar and therefore is omitted. \( \square \)

Instead of setting \( \mu_0 := \epsilon \) as in (25), it is recommended that \( \mu_0 \) is chosen conservatively if \( P \{\gamma_k = 1\} \) is not known exactly. In particular, we require

\[
\Delta \lambda \sum_{k=0}^\infty \beta^k E_k \{\|H x_{k}\|^2\} \leq \epsilon - \mu_0
\]

so that the bound (27) remains valid, which suggests that \( \mu_0 \) should be chosen as the minimal value such that

\[
\sum_{k=0}^\infty \beta^k E_k \{\|H x_{k}\|^2\} \leq \mu_0.
\]
problem (23) is initially feasible. Then the gap between $\epsilon$ and $\mu_0$ provides a robustness margin to allow for uncertainty in the arrival probability of sensor measurements.

9 Numerical examples

In this section, we run three sets of simulations: (A) demonstrates that the closed loop system satisfies the bounds in (27) and (28), and compares performance with the unconstrained optimal LQG controller; (B) shows that the averaged undiscounted closed loop cost accumulated over an infinite horizon is finite under the conditions of Theorem 9; and (C) demonstrates that the closed loop system satisfies (42).

Simulation A: We consider a system obtained by discretising a linearised continuous time model of a double inverted pendulum with a sample time of 0.01 s as in Davison (1990). The system matrices are

$$A = \begin{bmatrix} 1.0005 & 0.01 & -0.0005 & 0 \\ 0.998 & 1.005 & -0.0081 & -0.0005 \\ -0.0005 & -0.0005 & -0.0005 & 0.015 \\ -0.0081 & -0.0005 & 0.2942 & 1.0015 \end{bmatrix},$$

$$B = \begin{bmatrix} 0.0001 & -0.0001 \\ -0.0001 & 0.0003 \\ -0.02 & 0.05 \end{bmatrix}, \quad C = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \end{bmatrix}, \quad D = I,$$

and $\omega_k \sim N(0, \Sigma_w), \quad v_k \sim N(0, \Sigma_v), \lambda = 0.6.$ Here $\Sigma_w = \text{diag}\{0.5, 0.2, 0.9, 0.3\}$ and $\Sigma_v = 1.1I$. Initial conditions are given by

$$x_0 = \begin{bmatrix} -0.8 \\ 0.5 \\ 0.5 \end{bmatrix}, \quad \hat{x}_0 = \begin{bmatrix} 0.1 \\ 0.1 \\ 0.05 \end{bmatrix}, \quad \Sigma_0 = \begin{bmatrix} 0.5 & -0.5 & 0.5 \\ -0.5 & 0.5 & -0.5 \\ 0.5 & -0.5 & -0.5 \end{bmatrix}.$$

The constraint of (2) is defined by $\beta = 0.8, \epsilon = 2$ and

$$H = \begin{bmatrix} 0 & 0.1 & 0 \\ 0 & 0.1 & -0.1 \end{bmatrix}.$$

The weighting matrices in the cost function of (2) are given by $Q = \text{diag}\{10, 0.1, 10, 0.1\}, \quad R = 0.01I$. We choose a prediction horizon $N = 5$, as the unconstrained LQ-optimal, $K_{LQ}$, with respect to $(A, B, Q, R)$, and $M = \hat{\Sigma}C^T(C\hat{\Sigma}C^T + \Sigma_v)^{-1}$, where $\hat{\Sigma}$ is the solution of the algebraic Riccati equation

$$\hat{\Sigma} = A\Sigma A^T + \Sigma_w - \lambda A\Sigma C^T(C\Sigma C^T + \Sigma_v)^{-1}C\Sigma A^T.$$

For this system, cost and constraint, the solution of (23) yields $J_0 = 9.0757 \times 10^5$.

To verify (27) and (28) experimentally, we consider the average cost and constraint values over $10^8$ simulations, each of which has a run time of 150 time steps. For comparison, we run the same number of simulations with the same $\{\omega_k\}, \{v_k\}, \{\gamma_k\}$ sequences using the unconstrained optimal LQG controller, where $u_k = K_{LQ}^+k$ and the estimator gain is time-varying and given by $M = \Sigma_kC^T(C\Sigma_kC^T + \Sigma_v)^{-1}$. Here $\Sigma_k$ evolves as $\Sigma_{k+1} = \Sigma_k A^T + \Sigma_w - \gamma_k A\Sigma_k C^T(C\Sigma_k C^T + \Sigma_v)^{-1}C\Sigma_k A^T$. The results summarised in Table 1 agree with the bounds in (27) and (28) and show that, although the LQG controller gives a smaller closed loop cost (as expected), it violates the constraint. Note that $\beta^{150} = 2.9074 \times 10^{-15}$, so a further increase in the simulation run time has negligible effect on the cost and constraint estimates.

| empirical cost | MPC controller | LQG controller |
|----------------|----------------|----------------|
| $\beta_1$     | 0.98           | 0.99           |
| $\beta_2$     | 0.99           | 0.999          |
| $\text{tr}(Z_1X)$ | 518.3913       |

Table 1

Average discounted cost and constraint values for Simulation A

Simulation B: For this simulation the system model, cost and constraint parameters are the same as for Simulation A, except that the state matrix is redefined as $A = A - 1.175I$ and the unconstrained LQ-optimal feedback gain $K_{LQ}$, the steady state Kalman filter gain $M$ are modified accordingly. The modified matrices $A$ and $M$ satisfy (39) and there exists a scalar $\sigma = 4.2759$ such that (33) holds and $\sigma < \frac{1}{\gamma_2} = 5$. To estimate empirically the LHS of (36), we run $10^2$ simulations, each of which has a run time of $10^4$ time steps, and we thus obtain an estimate of the average cost value from the average over $10^5$ simulations as 14.5375. For this example therefore, the LHS of (36) is finite as implied by Theorem 9. Further increases in the simulation run time cause negligible changes in this estimate.

Simulation C: In this simulation we allow for different discount factors $\beta_1$ and $\beta_2 = \beta = 0.8$ in problem (2) and we set $\epsilon = 3.8$. The weighting matrix for control inputs is given by $R = 0.001I$ and the unconstrained LQ-optimal feedback gain, $K_{LQ}$, changes accordingly. All other model and problem parameters are the same as for Simulation A. We solve the steady state equation of (20) and obtain the RHS of (42) as 518.3913. To estimate empirically the LHS of (42) as $\beta_1$ approaches 1 while keeping $\beta_2$ fixed, we run three tests with $\beta_1$ equal to 0.98, 0.99 and 0.999, respectively. Each test consists of 100 simulations, each of which has a run time of $10^4$ time steps. Empirical values of the LHS of (42) corresponding to different values of $\beta_1$ are summarised in Table 2, and are in agreement with the bound (42). Note also that a further increase of the simulation run time only results in small variations in these estimates.

| empirical cost | Simulation C |
|----------------|--------------|
| $\beta_1 = 0.98$ | 483.3557     |
| $\beta_1 = 0.99$ | 472.8744     |
| $\beta_1 = 0.999$ | 468.7458     |
| $\text{tr}(Z_1X)$ | 518.3913     |

Table 2

Average undiscounted cost values for Simulation C

10 Conclusion

This paper proposes an output feedback MPC algorithm for linear discrete time systems with additive disturbances and noisy sensor measurements transmitted over
a packet-dropping communication channel. By parameterising a control policy in terms of affine functions of future observations, we provide a convex formulation of a stochastic quadratic regulation problem subject to a discounted expectation constraint. Our controller ensures recursive feasibility of the MPC optimisation problem and ensures constraint satisfaction and a discounted cost bound in closed loop operation. We provide a sufficient condition on the discount factor to ensure that the averaged undiscounted closed loop cost is finite. We consider conditions to ensure closed loop stability and investigate the effects of uncertainties in the probability of successfully receiving a sensor measurement.

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