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Processivity and Coupling in Messenger RNA Transcription

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**Abstract**

**Background:** The complexity of messenger RNA processing is now being uncovered by experimental techniques that are capable of detecting individual copies of mRNA in cells, and by quantitative real-time observations that reveal the kinetics. This processing is commonly modelled by permitting mRNA to be transcribed only when the promoter is in the on state. In this simple on/off model, the many processes involved in active transcription are represented by a single reaction. These processes include elongation, which has a minimum time for completion and processing that is not captured in the model. This simple on/off model, the many processes involved in active transcription are represented by a single reaction. These processes include elongation, which has a minimum time for completion and processing that is not captured in the model. This processing is commonly modelled by permitting mRNA to be transcribed only when the promoter is in the on state.

**Methodology:** In this paper, we explore the impact on the mRNA distribution of representing the elongation process in more detail. Consideration of the mechanisms of elongation leads to two alternative models of the coupling between the elongating polymerase and the state of the promoter: Processivity allows polymerases to complete elongation irrespective of the promoter state, whereas coupling requires the promoter to be active to produce a full-length transcript. We demonstrate that these alternatives have a significant impact on the predicted distributions. Models are simulated by the Gillespie algorithm, and the third and fourth moments of the resulting distribution are computed in order to characterise the length of the tail, and sharpness of the peak. By this methodology, we show that the moments provide a concise summary of the distribution, showing statistically-significant differences across much of the feasible parameter range.

**Conclusions:** We conclude that processivity is not fully consistent with the on/off model unless the probability of successfully completing elongation is low—as has been observed. The results also suggest that some form of coupling between the promoter and a rate-limiting step in transcription may explain the cell’s inability to maintain high mRNA levels at low noise—a prediction of the on/off model that has no supporting evidence.

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**Introduction**

The on/off model of gene activation accounts for the noise (standard deviation/mean) observed in mRNA and protein levels in single cells [1–7]. This model, shown in Figure 1, explains why cell to cell variations in mRNA may be greater than that of a simple Poisson process [2,4], that noise strength (variance/mean) varies with the rate of mRNA production [5], and predicts that the same population mean can be achieved with greater or lesser noise as determined by frequency of promoter activation [5]. The variation in mRNA distribution in vivo in yeast has recently been shown to uncover different modes of gene expression, with many yeast genes being less noisy than might be supposed [7]. Noise in protein levels has also been characterised as greater than Poissonian, and has been shown to be pathway specific with stress response genes being among the most noisy [8,9]. It has also been observed that protein distributions can have longer tails than would be the case if they followed a Poisson distribution, and that this may indicate the existence of generating processes other than a Poisson process [9]. Single molecule studies of E. coli proteins have demonstrated that in steady-state they follow a gamma distribution [10], while protein bursts follow a geometric distribution [11]. Noise, or noise strength, is a useful summary of variability, but does not capture any asymmetry in the distribution. Hence, measures of this asymmetry are of interest when analysing models and data.

Imaging experiments are the primary method for obtaining mRNA data at the copy per cell level, and while these are often relatively high-throughput techniques, the resulting histograms often show bin-to-bin variability. When characterising the shape of a distribution, skewness and kurtosis (see Materials and Methods for definitions) provide an alternative to methods that analyse the histogram, such as Kullback-Leibler (KL) divergence. To illustrate the advantage, Table S1 presents the results of a computational experiment where between 100 and 100,000 samples are drawn from a Poisson distribution and the skewness, kurtosis and KL divergence are calculated. This experiment indicates that skewness can be calculated reasonably accurately for sample sizes of 100
Proximity and Coupling

The On/Off Model of Transcription

The on/off model of transcription is represented as a Dizzy model and simulated using the Gillespie algorithm (see Materials and Methods). Figure 1 depicts the model, and a listing of it can be found in Appendix S1. This model specifies that mRNA (M) is synthesised at a rate of only when the gene is in the active state, and that mRNA is degraded at a rate  in any state where . The transitions of activation and inactivation have rates a and b respectively, and move the gene from the inactive (I) to the active (A) state (off to on) and vice-versa.

A useful distinction can be drawn between the rate of mRNA production when the gene is in the active state, and the process of transcription as a whole, which is the observable rate of production of mRNA over time, or in a population of cells, and incorporates the modulisation of  by the on/off switching rates a and b. We refer to this phenomenon as synthesis in the following discussion.

An analytic solution exists for the on/off model [28] and will serve as a reference for the simulation-based approach adopted here. Stochastic simulation is an alternative to solving complex sets of equations, and yields the probability distributions for species of molecules, in addition to the mean and variance that can be derived analytically. Computational simulation can also be used to determine the time course of a stochastic model.

The On/Off Model of Transcription

The on/off model was originally expressed by the following equations:

\[ I \xrightarrow{a} A \]  

Figure 1. The on/off model of gene activation and synthesis. Circles represent species or states, rectangles represent reactions between species or transitions between states. doi:10.1371/journal.pone.0008845.g001
\[ A \xrightarrow{b} I \] (2)

\[ A \xrightarrow{c} A + M \] (3)

\[ M \xrightarrow{d} \emptyset \] (4)

The steady-state variance has two terms: The first corresponds to the mean, and therefore the ratio of the variance to the mean will always be greater than 1. That is, the noise strength will be greater than that of a Poisson process which has a ratio of 1. It is noted in [4,28] that the model is normalised by the degradation rate \( d \). That is, the degradation rate sets the timescale over which the model evolves to steady state – the characteristics of the model are not otherwise dependent on \( d \). This feature is exploited in the following exploration of the model parameters.

A Survey of Parameters for the Normalised Model

Knowing that the model parameters can be rescaled to plausible values, the behaviour of the model is explored for a normalised degradation rate \( d = 1 \). It is helpful to structure the analysis by calculating \( b \) to give a desired value of \( M \) for selected values of \( a_1 \) and \( \epsilon_1 \), where the subscript 1 denotes normalised parameters. We choose \( M = 1, 2, 5, 10 \) and record characteristics of the distribution for the parameter combinations that yield the selected mean mRNA level. The survey is limited to low values of \( M \) as many constitutively expressed yeast genes have mRNA abundances in this range. The statistics of interest are the mean, variance, skewness, kurtosis, noise strength and the histogram of the simulated distribution of \( M \). To condense the presentation of the results, we focus on mean values of \( M = 1 \) and 10.

The values of \( a_1 \) we survey range from 1/8 to 128, and the values of \( \epsilon_1 \) range from 10 to 5000. (The parameter values specify the reaction probability density per unit time.) These ranges are selected to include values determined by [4] to characterise noise in mammalian mRNA (i.e. \( a_1 \) around 1 and \( \epsilon_1 \) between 500 and 10000). Values of \( a_1, \epsilon_1 \) as low as 10 are explored because values of 500 and above may not give plausible synthesis rates on rescaling (this issue is discussed below). Higher values of \( a_1 \) are explored in order to discover the behaviour of a system that transitions more rapidly from off to on, with respect to the degradation rate. Often, \( b > a \) and in most cases \( b \gg a \). This means that the time spent in the inactive state \((1/a)\) is typically much greater than the time spent in the active state \((1/b)\). The ratio \( a_1/(a_1 + b_1) \) is also of interest as it represents the fraction of time spent in the on state. This ratio is in fact constant for each series of values of \( a_1 \) explored here, as \( M \) is held constant, and by equation 5: \( a_1/(a_1 + b_1) = M/\epsilon_1 \). We show that varying \( a_1 \) while keeping \( M \) and \( \epsilon_1 \) (and therefore this ratio) constant results in a wide range of behaviour.

The analysis presented here is organised around \( a_1 \). When \( a_1 = 1 \) the time in the inactive state equals that taken for a molecule of \( M \) to degrade, on average. It should be noted that the two modes of gene expression \( b > c \) and \( b < c \) identified in [7] correspond to \( a_1 > 1 \) and \( a_1 < 1 \) for \( M = 1 \), and to \( a_1 > 10 \) \( a_1 < 10 \) for \( M = 10 \) in the analysis presented here. Modes of expression are addressed in more detail later, where we also show that the range of \( a_1 \) values surveyed includes all those we could obtain from the literature. As we are primarily concerned with noise in the transcription process itself (intrinsic noise), noise strength is taken as the most appropriate measure of variability. It should be noted that noise strength scales with the mean, which can lead to unintended artifacts where both intrinsic and extrinsic noise is considered [29].

Simulation Results

Figure 2 (A and C) shows that the simulated values for mean (plotted as open symbols) and variance (plotted as filled symbols) lie on the curves predicted by equations 5 and 6, as would be expected. Note that the color coding used in this plot, and those following, identifies a series of results according to the value of synthesis \( a_1 \), e.g. orange triangles denote models where \( a_1 = 50 \), the independent variable (activation) is on the x axis and the dependent variables on the y axis. Variance is proportionately higher when \( M = 10 \) in comparison with \( M = 1 \). In both cases, for activation rates \( a_1 \geq 10 \), the mean and variance tend towards the same value, and hence noise strength tends to 1. Figure 2 (B and D) shows that skewness and kurtosis reduce as \( a_1 \) increases from its lowest value, and that the range of skewness and kurtosis values is not strongly dependent on the mean value of \( M \). The variation of skewness and kurtosis with \( a_1 \) is highest for the largest synthesis rates, \( a_1 \), and this reduces considerably for low synthesis rates. Skewness and kurtosis reduce to approximately 1 for a mean \( M = 1 \) (2B), while for \( M = 10 \) (2D) they reduce to 0.34 and 0.13 respectively. These results confirm that the second, third and fourth moments of the mRNA distribution are only weakly dependent on \( a_1 \) when \( a_1 \) is above 2. The error bars for skewness and kurtosis are plotted in Figure S1.

Figure 3 shows the extent to which the on/off model can be fitted by gamma, Poisson, Gaussian and negative binomial distributions. Neither the gamma nor Gaussian model fits as well as the negative binomial distribution. The Poisson distribution fits well for certain values of activation, depending on the mean \( M \). With the exception of \( \epsilon_1 \) values of 1000 and below (\( M = 10 \)), and consistent with the analysis of [30], the on/off distribution is accurately modelled by the negative binomial distribution.

These results suggest the mRNA distribution can never be Gaussian for low mean values, while it can take a more symmetric form as \( M \) and \( a_1 \) increase. The results account for both the behaviour shown in [3] where the parameters give a Poisson or normal-looking distribution, and that described in [4] where a much noisier parameterization is chosen. A consequence of these findings is that fitting parameters in the space where \( a_1 \gg 1 \) will be hampered by the lack of variation of the distribution with \( \epsilon_1 \). The error bars on skewness and kurtosis show that we could not confidently determine activation rates from these values, although they do vary with \( a_1 \).
Parameter Interpretation

As noted, the degradation parameter \(d\) specifies the timescale. Letting \(a_s, b_s\) and \(c_s\) be the parameters for some other time scale \(s\) as determined by the degradation rate \(d_s\), then the scaled model is equivalent to the normalised model providing:

\[
a_s = a_1 \times d_s, \quad b_s = b_1 \times d_s \quad \text{and} \quad c_s = c_1 \times d_s.
\]

That is, the predictions for the normalised parameters can be rescaled to determine the absolute rates (or absolute mean time for a transition) given knowledge of \(d_s\). The plausibility of the off-on activation rate can then be assessed, as can the feasibility of the synthesis rate \(c_s\). The scaled values for \(a_s\) for degradation rates of 1s, 60s, 10min, 1 hour and 4 hours are given in Table S2. The scaled \(c_s\) values are given in Table S3.

From considerations of the initiation rate (0.2), physical size (100 bp) and average elongation rate of Pol II (20–30 bp/s), it can...
be concluded that an mRNA transcript can be completed at a maximum rate of one every 5 seconds, i.e. a rate of 0.2 [31] (assuming also that events in elongation, splicing and termination apply uniformly to transcripts). Allowing for the possibility of a higher maximum elongation rate of up to 70 bp/s [26], we would limit the maximum mRNA production rate $c_1$ to lie in the range 0.2–0.5. Although inexact, such an estimate is useful as it places constraints on plausible $c_1$ values for a known $d_0$. Table S3 lists the scaled values (i.e. the product of the normalised $c_1$ rates and the actual degradation rate). Degradation rates observed for some yeast mRNAs are in the order of minutes while in mammalian systems the rate may be in the order of hours. Taking the half life of a yeast mRNA as 10 minutes, values of $c_1$ of 500 are not plausible as they correspond to mRNA being transcribed above the maximum rate. This value for $c_1$ is plausible for mammalian mRNA synthesis where the degradation rates can be 4 hours [4]. By considering how the normalised parameters are scaled, we can constrain the plausible range for model parameters. Naturally, our beliefs about plausible rates may change as new evidence emerges.

Once the parameters have been rescaled, the average times a gene spends in the active and inactive states can be worked out - we can consider whether it is plausible for a gene be active for 1/10s, or for 100s. This requires biological insight and a deeper model of transcription that incorporates events at the molecular level.

Models of Elongation

The production of a mature eukaryotic messenger RNA involves many processing events that can occur co-transcriptionally, including 5’ end capping, intron removal (splicing) and 3’ end cleavage and polyadenylation [14]. Some of these events can now be observed and measured at the single-transcript level. The extent to which these transcription and processing events are coupled is still under investigation and stochastic modelling may play a role in formulating and testing theories. However, the on/off model clearly does not account for the many hundreds of molecular events that occur in the production of a transcript, nor of the abortive transcriptional activity that is now being uncovered. In this section, we examine several modifications that change the characteristic of the synthesis step by representing the elongation process in more detail. It should be noted that elongation is not typically considered the rate limiting step in synthesis, the prior initiation step will set the rate for the synthesis of mature transcripts. However, under high induction the spacing of Pol II sets an upper limit as discussed above. And if, in addition, the gene is short, we can consider the case where elongation does become the sole and limiting factor.

A model of elongation might represent every base (or nucleotide) addition event for each of the N bases in the transcript. Figure 4A shows how the elongation process might be incorporated into the on/off model. Rather than fix a value for N to model a specific gene, we note that, for any large N, synthesis is now modelled as a more deterministic process [19,25]. A parameterization for the model is obtained by giving each of the N transitions a rate N times that of the original synthesis step of the on/off model. The key questions are how the mRNA distribution is affected by modifying the model in this way, whether these changes are reflected in observable differences in the mean, variance, skewness and kurtosis, and whether any such changes can be used in model fitting and model selection.

In exploring the parameter space, we make the simplifying assumption that at most one Pol II is active at any time, and hence ignore any effects of polymerase assemblies holding each other up during elongation (as noted earlier, this phenomena has been investigated by [19,20,23]). This assumption is borne out for some yeast genes where typically 0 or 1 nascent mRNAs are detected, for example MDN1 [7], (while an average of two polymerases were observed at the transcription site in [26]), but clearly this assumption will not apply to longer genes expressed at high levels where multiple Pol IIs transcribe mRNA simultaneously. If the gene is short, and the induction level is high, then we may ignore the time between the completion of each round of transcription and effectively consider elongation as the rate setting step. Were Pol II to be recycled at the promoter, we would have exactly this case. This simplification mirrors that in the original on/off model
where the initiation rate and the time for synthesis cannot be distinguished from the synthesis rate.

In the processive and coupled elongation models, the steps in elongation are represented by a counting procedure rather than by explicitly creating a series of N transitions. This is preferable as simulating models containing long series of transitions is more computationally expensive than the counting approach, and the models they implement are equivalent. The elongation transition increments a count place CT. When the count reaches a specified value the number of mRNA $M$ is increased by 1 (as indicated by the filled rectangle that represents the testing of the count value, see Figure 4B). Letting $N = 100$, the synthesis rate is multiplied by 100 to obtain the elongation rate $e$. When the gene is active, the elongation process takes the same amount of time as the original synthesis step, and so the rate of mRNA production should be the same as in the original model. But the variability of the synthesis time is much reduced, and so the effect on the distribution characteristics must be investigated. The elongation process is modelled stochastically, but has a more predictable outcome. A narrow distribution of elongation times also follows from similar assumptions made in [19]. As the elongation rate is easily derived from the $c_1$ parameter, it is straightforward to assess the behaviour of the new model using the same $a_1$ and $b_1$ parameters as before. No analytic solutions are available for the models described below; however, it is possible to show the effects of the modifications by comparison with simulations of the on/off model (or the analytic solutions for it).

Processive Elongation - the On/Off-PE Model

In a simple deterministic elongation model, the on/off/PE model would be retained as all elongation events contribute to $M$ (see the Text 1 for more details of this model). However, when the count is interpreted as corresponding to the position of a Pol II on the gene, it is implausible that the count (Pol II location) is retained from one round of activation to the next. In fact, elongation is often considered to be processive, that is, Pol II assemblies will continue to transcribe mRNA even if the promoter becomes inactive. The on/off-PE model allows the count, once started, to continue until $M$ is incremented once, irrespective of the gene being active or inactive. This modification may allow an additional mRNA to be produced on each round of transcription. And, as the limiting reaction rate is greater than in the on/off model, this effect will be magnified at higher activation rates. (This magnification would be less were initiation an explicit step in the process: The processive model represents the upper limit for an initial rate-limiting step in synthesis.)

Coupled Elongation - the On/Off-CE Model

Another possibility is for Pol II s that are part-way through elongation when the gene transitions to the inactive state to disassociate from the DNA. (The inefficiency of engagement and elongation has been described in [26].) The on/off-CE model modifies the deterministic elongation on/off model by resetting the count when the gene becomes inactive. This change removes the contribution of previous elongation activity when the gene re-activates, and implements a simple coupling between the completion of elongation (which we equate with the production of a mRNA) and the state of the promoter. Elongation can be ‘abortive’ - elongation may not lead to a mRNA - and its successful completion is correlated directly, or indirectly, with the promoter.

Simulation Results

Beginning with the results for mean and variance for the on/off-PE model, Figure 5A shows that for a desired $M = 1$, the simulated mean rises above the desired value for all but the lowest values of $a_1$. For $a_1 = 2$, the mean is approximately double that intended. The mean continues to rise as $a_1$ increases, the range of values for the steady-state mean lies in the range 1.05–55.5 copies/cell. The variance takes the predicted values for $a_1 \leq 1$, but thereafter increases in parallel with the increase in the mean. The noise strength does not decrease below 0.9 except for the combination of $a_1 \geq 4$ and $c_1 \leq 100$. When $M = 10$ (Figure 5B), the mean values for $M$ predicted by the on/off-PE model begin to rise above the desired value of 10 as $a_1$ increases beyond 8. Mean mRNA lies between 10.1 and 111 copies/cell across the parameters surveyed. However, noise strength only reduces significantly below 1 for the combination of high values of $a_1$ and low values of $c_1$, as the increase in the mean is again matched by increasing variance.

Figure 5 (C and D) shows that for the on/off-CE model when $M = 1$, mean and variance both fall significantly below their predicted values as $a_1$ increases above 2. This is due to the increasing probability that the gene will switch to the off state while elongation is in progress, a consequence of which is the reduced overall rate of mRNA production. Noise strength tends to 0 as $c_1$ increases. Despite mean and variance diverging from the values predicted by the on/off model, they converge, and so the noise strength characteristic does not appear different from that of the on/off model.

The skewness and kurtosis of the on/off-PE model are compared to those of the on/off model in the scatterplot in Figure 6 A and B for $M = 1$ (and in Figure S2 for $M = 10$). The on/off-PE model has correlated, but reduced, skewness and kurtosis (with reference to the on/off model). As can be seen, the predictions differ by 2 or 3 standard deviations over much of the parameter space. The on/off-CE model shows a different characteristic, as indicated in Figure 6 C and D, where both measures vary non-monotonically, reducing as $a_1$ increases from $1/8$ to 2, then increasing again as the $M$ distribution becomes
increasingly concentrated around 0. For the on/off-CE model, when $M = 1$, skewness has a minimum value of approximately 2, and kurtosis remains above 4; and when $M = 10$, skewness has a minimum value of 0.54, and kurtosis remains above 0.34.

In summary, the models considered here have very similar characteristics when $a_1$ is less than 2, and $M$ is high (i.e. 10). It is most evident that different mechanisms are in play for low values of $M$. Differences are also observed when $a_1$ is greater than 2: It has been shown that using an elongation model as the representation of synthesis in the on/off model can explain values of noise strength below 1. The range of skewness and kurtosis values is shown to be largely independent of the mean, hence re-scaling is not necessary in order to compare values across datasets or models (i.e. in contrast with the variance, factoring by the mean is not necessary). Skewness values show differences of several standard deviations when comparing the deterministic, -PE and -CE models with the reference model, and this may be a general phenomena that is useful in model fitting and model selection.

The mechanisms proposed in the elongation models can be reviewed in the light of the simulation results. For more frequent activation/inactivation cycles, the processivity mechanism increases mRNA production significantly. The high probability of elongation initiating, with no means for elongation to pause or

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**Figure 5. Mean (open symbols) and variance (filled symbols) for the on/off-PE model.**

A. $\langle M \rangle = 1$; B. $\langle M \rangle = 10$. Points are average values from 10 repetitions. Error bars show the standard deviation for 10 repetitions. Solid lines are the theoretical solutions for mean and variance.

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for Pol II dissociate, is the reason for this outcome. Were the probability of the (single) Pol II completing elongation to be 1/90, as observed in [26], then the mean mRNA would be much reduced. A reduction of this scale cannot be applied when \( a_1 \) is low without stopping elongation, but could be a factor that reconciles the processive model with the basic on/off model for high values of \( a_1 \). In contrast, the coupling mechanism causes mRNA levels to be depressed. This suggests that for higher activation rates (\( a_1 \geq 8 \)), any coupling between the promoter and a rate limiting step in synthesis will prevent high mean mRNA levels being achieved at low noise. This effect is observed for the particular case where elongation is rate limiting, but would also follow if the escape of Pol II from the promoter-proximal pause region was linked to promoter state.

Modelling the Single-Cell

Two questions are now addressed in the analysis of single-cell data: i. do skewness and kurtosis measured from real data match values predicted from the model? and ii. is there evidence of constraints on the parameters of the on/off model when the model is fitted to single cell data?

We created a yeast strain that expresses a reporter RNA called Ribo1, and the number of mRNA per cell was counted using an in situ hybridization technique sufficiently sensitive to reliably detect single mRNA molecules (see Figure 7 and Materials and Methods). The reporter was under the control of a Tet-responsive promoter, contained the ACT1 intron and the 3' processing signal of PGK1. Four cultures of yeast strains expressing Ribo1 were prepared and analyzed independently. In a representative sample, the mean
number of mRNA/cell in steady-state was 23.8 copies/cell in a population of cells (327 cells, 2–82 copies/cell) The observed probability distribution is shown in Figure 8A, along with the predicted probability distribution under the parameterization listed in Table S4 (sample 4). As indicated in Table 1, skewness and kurtosis are in good accord with observation when the variability of the data is accounted for by estimating the standard deviation by the bootstrap method. At these elevated mean mRNA levels, kurtosis becomes negative, i.e. the distribution is flatter than Gaussian. The fit of these characteristics to the data could be improved if skewness and kurtosis were optimized in model fitting.

To explore the possibility of constraints on model parameters, the Ribo1 data are included in the scatterplots in Figure 8 of \( a_1 \)

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**Figure 7. Image-based measurement of mRNA in yeast.**

A. Schematic of the Ribo1 reporter gene used in this study. The position of the probes used to detect the reporter RNA are indicated by horizontal red bars. B. Detection of single molecules of the Ribo1 reporter RNA. Expressing (upper panels) and control cells (lower panels) are shown. Left panels display the RNA signal (red) overlayed with the nuclei (blue). Each field in a projection of a 3D stack (18×18×6 μm). C. Efficiency of RNA detection. Histograms of the number of spots versus spot intensities across an entire 3D stack (63×63×6 μm) are shown for control or Ribo1 expressing cells. The area shaded in yellow correspond to the spots included in the analysis after thresholding. Bottom panels: overlay of the two histograms revealing the amount of Ribo1 RNA molecules lost by the thresholding procedure (<15%; red area).

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**Figure 8. Modelling mRNA expression in single cells.**

A. Distribution of mRNA, a. observation and b. model. B. Scatterplot of \( c_1 \) and \( a_1 \) parameters from published models: blue triangles indicate mammalian data [4], green circles indicate yeast [7], as do purple squares (this paper), and the black diamond shows data from E. coli [2]. Closed symbols are plotted when \( b_1 > a_1 \), open symbols indicate \( b_1 < a_1 \). All parameters have been scaled such that \( d = 1 \) as in preceding figures. The line is an empirically-derived equation for the upper limit of \( c_1 \). C. Scatterplot of \( \langle M \rangle \) values and \( a_1 \) parameters from published models, symbols are defined as in B. Blue and black lines are the predicted upper constraint on \( \langle M \rangle \) for alternative assumptions about the proportion of time in the on state. The grey line is \( \langle M \rangle = a_1 \).

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values from published models [2,4,7]. Values of synthesis \( c_1 \) are plotted against \( a_1 \) in 8B, and in 8C mean mRNA \( \langle M \rangle \) is plotted against activation rate. This analysis shows that, on the available evidence, there is a constraint on the rate of synthesis that can be approximated by: \( c_1 < (10 - a_1/2)^2 \). When this constraint is incorporated into equation 5, a constraint on \( \langle M \rangle \) can be inferred when the ratio \( a_1/(a_1 + b_1) \) is considered. As this ratio increases, the possibility of high activation rates producing high levels of mRNA at low noise is reduced. The on/off model can be parameterized to give high copies/cell of mRNA with high activation rates (i.e. fast activation/inactivation cycles that reduce noise) but such a combination is not supported by the available data. Rather, high mean levels are modelled by low activation rates and are noisy. Low activation times can also explain data where the mean mRNA is low. This evidence further supports the argument that regulation acts to increase time in the active state or to effectively increase the synthesis rate by regulating Pol II activation to inactivation is another simple distinction that can be possibly held, above it a second mode is possible. The ratio of synthesis to inactivation has been suggested as a way to define kinetic mechanisms [5], and to determine transcriptional bursting, i.e. \( c/b > 1 \) or \( c/b < 1 \). Some limitations of this approach have been noted [7]. The ratio of synthesis to activation might similarly be taken as defining a mode of gene expression: \( c/a > 1 \) or \( c/a < 1 \). However, it is a straightforward consequence of equation 5 that for \( a_1 < \langle M \rangle \) it must be the case that \( c/b > 1 \) and \( c/a > 1 \). The boundary \( a_1 = \langle M \rangle \) is shown in Figure 8C: Below this line only one of the inequalities in each pair can possibly hold, above it a second mode is possible. The ratio of activation to inactivation is another simple distinction that can be made, and this too is shown in Figure 8C by the use of open and closed symbols. It can be seen that \( b_1 < a_1 \) only for higher values of \( a_1 \). The ratios \( b/a, c/a \) and \( c/b \) take a continuum of values that, broadly speaking, reduce as \( a_1 \) increases. None of these distinctions explain the absence of high \( a_1 \) rates for high \( \langle M \rangle \), but we note that the assumption of coupled of elongation reveals that this would be, at the very least, a highly inefficient mode of transcription under such conditions, if not an impossibility.

**Discussion**

The on/off model of mRNA transcription has been simulated over a wide range of feasible parameters. The distribution of mRNA/cell in steady-state has been summarised by the mean, second, third and fourth moments. Unlike the mean and variance, equations for skewness, kurtosis have not been derived analytically in the general case - and their application to the on/off model is presented here for the first time. Modifications to the on/off model that give a more detailed account of elongation are shown to influence all of these measures. Moreover, we have shown that skewness and kurtosis are of value when comparing models.

Each step in elongation, and the possibility of Pol II backtracking, is modelled in [19]. The count procedure in the models proposed here similarly captures the phenomenon of a repeated step-wise process, but does so in a way that allows the parameters of the on/off initiation model to be mapped to the new models, and surveyed. The representation of promoter state and elongation in a single model is also found in the model of the dynamics of Pol II engagement proposed by [26]. However, our model does not distinguish initiation from the later engagement process, nor assign them different rates in order to explore the assumptions we make in a uniform way across the parameter space.

The on/off model has been used successfully to explain single-cell mRNA data [1–7], without the modifications proposed here. Even where noise is observed to be low [7], it remains at least Poissonian and hence consistent with the model. While a deterministic elongation model reduces noise below this bound, processivity and coupling restore noise to the gene expression characteristic. Initiation, pausing and backtracking in elongation further reduce its predictability, and accounting for them in a processive model will add additional noise to the process of transcription. The marginal impact that different stochastic processes have on the overall size of fluctuations in a population is noted in [32], where protein noise is shown to be only weakly dependent on burst size and waiting time statistics, and on the number of steps in the degradation pathway. Through this coarse-graining effect [32], the phenomenological on/off model may indeed accurately summarise the molecular mechanisms that are being uncovered and modelled. One critical step will be to accurately determine the activation and synthesis rates experimentally, e.g. as in [2], as without this data they must be found by optimization. As illustrated in this paper, the distributions that they are estimated from are essentially invariant in certain regions of parameter space. A further issue arises in model comparison, where the saturated elongation process necessary to maintain a high mRNA expression level may obscure any characteristics that result from molecular-level elongation mechanisms: Even with full knowledge of the important parameters, the predictions are likely to be dominated by the activation/inactivation cycle rather than by the more subtle features of the synthesis process at lower levels of processing. From the modelling perspective, genes expressed at low levels are the more promising candidates for testing alternative theories, but clearly this poses challenges experimentally.
Materials and Methods

Stochastic Simulation

All simulations were performed using the Gillespie algorithm implemented in Dizzy (http://magnet.systemsbiology.net/software/Dizzy/). This software was modified to store the histograms of the distributions for further analysis. 10,000 simulations were run for each set of parameters surveyed, and this procedure was repeated 10 times to obtain means and variances. The EDDIE compute cluster was used to run the simulations, which, for the more complex models, took 7–10 hours for each parameter survey, each of which was repeated 10 times for the four models.

It is necessary to show that the simulations have stabilised, and this was achieved by measuring the divergence between the distribution for the end time and that of an earlier time using the Kullback-Leibler (KL) measure:

\[ KL(p,q) = \sum_{x} p(x)\log(p(x)/q(x)) \] (8)

The KL divergence is calculated from the probability that the data falls into a bin \( \hat{x} \) as defined in the two distributions being compared, \( p \) and \( q \). The divergence between the probability distribution of mRNA at 50s and at 33s (of simulated time) is plotted in Figure 3 in order to demonstrate that the simulation has reached a stationary state. KL divergence is also useful for determining how good a fit the best-fitting Poisson, Gaussian, gamma or negative binomial distribution is to the simulated distribution. The best-fit distribution parameters are calculated from the simulation data using maximum likelihood, but this calculation alone does not return a goodness of fit, and so the distance must be calculated separately. The SSJ Java library (http://www.iro.umontreal.ca/~simardr/ssj/indexe.html) was used to obtain the distribution parameters for both cases: A. skewness; B. kurtosis. Points are average values from 10 repetitions. Solid grey line indicates \( y = x \).

Kurtosis and Kullback-Leibler divergence from sampling a Poisson distribution (\( \lambda = 9 \)). Mean and variance are obtained from 1000 repetitions. This table presents the results of a computational experiment where between 100 and 100,000 samples are drawn from a Poisson distribution and the skewness, kurtosis and KL divergence are calculated. The value of \( \lambda \) is 9 in the generating distribution, and therefore the theoretical value of skewness is \( 1/3 \), kurtosis is \( 1/9 \), and the theoretical mean is \( 9 \) in all samples. The table gives the mean and variance for these measures based on 1000 repetitions, and lists the error. For skewness and kurtosis, the error is the magnitude of the difference between the mean and the theoretical value, and for KL divergence the error is the factor by which the divergence increases in comparison with the divergence for the larger sample size (i.e., that given in the row above). This exploration indicates that skewness can be calculated reasonably accurately for sample sizes of 100 and above, while kurtosis may require a greater number, possibly as many as 1000 samples. For KL divergence, the difference between a sampled distribution and the theoretical
distribution (which is known in this case) increases by a factor of approximately 8 for a ten-fold reduction in the number of samples. Therefore, skewness and kurtosis appear to be robust measures with respect to sample size, while KL divergence suffers from the bin-to-bin variability that results from lower sample sizes.

Table S2: Values of $a_d$ scaled for $d = 1s, 1min, 10min, 60min$ and expressed as mean times (s or h:m:s) for the transition. Found at: doi:10.1371/journal.pone.0008845.s003 (0.02 MB PDF)

Table S3: Values of $c_d$ scaled for $d = 1s, 1min, 10min, 60min$ and expressed as mean times (s or h:m:s) for the transition. Found at: doi:10.1371/journal.pone.0008845.s004 (0.04 MB PDF)

Table S4: Parameters for the on/off model for four sample populations of Ribo1.

References

1. Becskei A, Sauermann B, van Oudenaarden A (2003) Contributions of low molecule number and chromosomal positioning to stochastic gene expression. Nat Genet 35: 535–539.

2. Golding I, Paulsson J, Zawliski S, Cox EC (2005) Real-time kinetics of gene activity in individual bacteria. Cell 123: 1025–1036.

3. Goss P, Pecou J (1998) Quantitative modelling of stochastic systems in molecular biology by using stochastic Petri nets. PNAS 95: 6750–6755.

4. Raj R, Peskin C, Trachina D, Vargas D, Tyagi S (2006) Stochastic mRNA synthesis in mammalian cells. PLoS Biology 4: e309.

5. Raser J, O’Shea E (2004) Control of Stochasticity in Eukaryotic Gene Expression. Science 304: 1142–1145.

6. Volotin D, Maricnik J, Blake W, Ostoff N, Tsimring L, et al. (2006) Origins of extrinsic variability in eukaryotic gene expression. Nature 439: 861–864.

7. Zengelus D, Larson D, Singer R (2006) Single-RCN counting reveals alternative modes of gene expression in yeast. Nat Struct Mol Biol 13: 1263–1267.

8. Bar-Evan A, Paulsson J, Maheshri N, Carmi M, O’Shea E, et al. (2006) Noise in protein expression scales with natural protein abundance. Nature Genetics 38: 636–643.

9. Newman J, Ghannamghami S, Ihmels J, Brosh D, Noble M, et al. (2006) Single-cell proteomic analysis of s. cerevisiae reveals the architecture of biological noise. Nature 441: 840–846.

10. Cai L, Friedman N, Xie XS (2006) Stochastic protein expression in individual cells at the single molecular level. Nature 440: 389–393.

11. Yu J, Xiao J, Ren X, Luo K, Xie XS (2006) Probing Gene Expression in Live Cells, One Protein Molecule at a Time. Science 311: 1600–1603.

12. Blake W, Kaern M, Cantor C, Collins J (2003) Noise in eukaryotic gene expression. Nature 422: 633–637.

13. Blake W, Kaern M, Cantor C, Collins J (2005) Rules of engagement: co-transcriptional recruitment of pre-mRNA processing factors. Current Opinion in Cell Biology 17: 251–256.

14. Hahn S (2007) Structure and mechanism of the RNA polymerase II transcription machinery. Nature Structural and Molecular Biology 11: 394–403.

15. Fuda NJ, Ardehali MB, Lis JT (2009) Defining mechanisms that regulate rna polymerase ii transcription in vivo. Nature 461: 186–192.

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Appendix S1: On/off models in Dizzy syntax. Found at: doi:10.1371/journal.pone.0008845.s007 (0.03 MB PDF)

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Author Contributions

Conceived and designed the experiments: SA. Performed the experiments: SA MCR EB. Analyzed the data: SA. Contributed reagents/materials/analysis tools: MCR RDA IG EB JDB. Wrote the paper: SA.