Prognostic value of high-sensitivity cardiac troponin I in heart failure patients with mid-range and reduced ejection fraction

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Abstract

Background
The identification of high-risk heart failure (HF) patients makes it possible to intensify their treatment. Our aim was to determine the prognostic value of a newly developed, high-sensitivity troponin I assay (Atellica®, Siemens Healthcare Diagnostics) for patients with HF with reduced ejection fraction (HFrEF; LVEF < 40%) and HF with mid-range EF (HFrEF) (LVEF 40%–49%).

Methods and results
A total of 520 patients with HFrEF and HFrEF were enrolled in this study. Two-year all-cause mortality, heart transplantation, and/or left ventricular assist device implantation were defined as the primary endpoints (EP). A logistic regression analysis was used for the identification of predictors and development of multivariable models. The EP occurred in 14% of the patients, and these patients had higher NT-proBNP (1,950 vs. 518 ng/l; p < 0.001) and hs-cTnI (34 vs. 17 ng/l, p < 0.001) levels. C-statistics demonstrated that the optimal cut-off value for the hs-cTnI level was 17 ng/l (AUC 0.658, p < 0.001). Described by the AUC, the discriminatory power of the multivariable model (NYHA > II, NT-proBNP, hs-cTnI and urea) was 0.823 (p < 0.001). Including heart failure hospitalization as the component of the combined secondary endpoint leads to a diminished predictive power of increased hs-cTnI.
Conclusion
hs-cTnI levels $\geq 17$ ng/l represent an independent increased risk of an adverse prognosis for patients with HFrEF and HFmrEF. Determining a patient’s hs-cTnI level adds prognostic value to NT-proBNP and clinical parameters.

Introduction
The prognosis of patients with chronic heart failure is rather poor; the 3-year all-cause mortality is approximately 35% [1]. Various prognostic scoring systems can identify the highest-risk patients and warn their physicians that, in accordance with clinical practice guidelines, further diagnostic tests need to be done, pharmacotherapies should be modified, non-pharmacological treatments should be considered, or, as a last resort, left ventricular assist device (LVAD) implantation or heart transplantation (HTX) should be considered [2]. There are a limited number of models for stable patients with chronic heart failure that evaluate cohorts of patients treated according to the current guidelines and, apart from clinical and classic laboratory parameters, also use natriuretic peptides and other novel biomarkers. In addition, most models are based on either overall mortality or cardiovascular mortality as the monitored endpoint, thus omitting a particularly important component of terminal heart failure treatment, i.e. LVAD implantation or heart transplantation, both of which can alter the natural course of the disease and improve the patient’s prognosis. Importantly, the discriminatory power of the currently used models tends to be lower when monitoring hospitalisations for heart failure [3]. For these reasons, we used the overall mortality, LVAD implantation, and/or HTX as the primary combined endpoints in our study. The secondary combined endpoints were defined as combination of primary endpoints and acute heart failure hospitalization.

Establishing the natriuretic peptide levels is integral to providing care for heart failure patients, as these peptides are the most important prognostic markers [4, 5]. High-sensitivity cardiac troponin is another cardiac biomarker that is readily available online and in standard clinical laboratories. A recently published meta-analysis showed that hs-cTnT is an independent predictor of heart failure patients’ prognoses [6]; however, there is only a limited amount of information on the prognostic significance of hs-cTnI.

The aim of our analysis was to describe the prognostic significance of a novel high-sensitivity cardiac troponin I assay (Atellica® IM high-sensitivity troponin I, Siemens Healthineers) in a cohort of patients with stable systolic chronic heart failure (reduced or mid-range ejection fraction) and evaluate its contribution to NT-proBNP levels and clinical parameters.

Methods
This study protocol complies with the regulations set forth in the Declaration of Helsinki and was approved by the Ethics Committee of the University Hospital Brno (Brno, Czech Republic). Written informed consent was obtained from each of the patients before they began their participation in the study. The study protocol was described in our previous study [4]; in short, a total of 1,088 patients were prospectively recruited from November 2014 to November 2015, and 520 of them were evaluated in the current sub-study. The primary endpoint, i.e., the two-year prognosis for all-cause mortality, heart transplantation, and/or left ventricular assist device (LVAD) implantation, was evaluated up to November 2017. The secondary combined endpoints were defined as combination of the two-year all-cause mortality, heart
transplantation, LVAD implantation and/or acute heart failure hospitalization. The patients were followed up prospectively at outpatient departments, and the mortality rates were verified using the centralised database of the Czech Republic Ministry of Health. The monitored data from the patients were gathered at the end of the two-year follow-up period.

**Study population**

Patients with a stable form of either heart failure with reduced ejection fraction (HFrEF) (LVEF < 40%) or heart failure with mid-range ejection fraction (HFmrEF) (LVEF 40%–49%) were eligible for inclusion in this study. The cohort included patients who were followed up and treated in outpatient clinics of cardiology departments (where authors of this study work) for stable chronic heart failure. All three cardiology departments provide specialised care for heart failure patients. In their medical history, all patients included in the study had an attack of heart failure with elevated natriuretic peptides and a reaction to heart failure treatment. Echocardiography was performed for all patients by experienced physicians working in echocardiography laboratories of cardiology departments. Other structural and/or functional abnormalities related to the patients’ heart failure were found, including LVEF values ranging between 40% and 49%: left ventricular hypertrophy (an interventricular septum ≥ 11 mm or a left ventricular mass index ≥ 115 g/m^2^ for the men and ≥ 95 g/m^2^ for the women), left atrial enlargement (a left atrial volume index > 34 ml/m^2^) and/or diastolic dysfunction (E/e’ ≥ 13 and mean e’ < 9 cm/s). A current NT-proBNP level < 125 pg/mL was not among the exclusion criteria because all the patients had increased NT-proBNP levels in their medical history. The exclusion criteria were the following: not signing the informed consent, signs and symptoms of acute decompensation from heart failure, and conditions other than heart failure that would likely hinder the patients’ mid-term prognosis (e.g., advanced cancer, severe dementia, etc.). The final decision about the diagnosis of chronic heart failure and the enrolment of a specific patient in the study was done by a cardiologist experienced in care for heart failure patients.

**Laboratory methods**

The patient plasma NT-ProBNP levels were analysed using the Cobas E411 NT-proBNP electrochemiluminescence immunoassay Kit (Elecsys proBNP II, Roche Diagnostics, Indianapolis, IN, USA). The limit of blank (LoB) was 3 pg/mL, the limit of detection (LoD) was 5 pg/mL, the measuring range was 5–35,000 pg/mL, the functional sensitivity (the lowest analyte concentration that can be reproducibly measured with an intermediate precision CV of 20%) was 50 pg/mL, and the cut-off value was 125 pg/mL.

The plasma hs-cTnI levels were analysed using the Atellica® IM High sensitivity Troponin I assay (Atellica® IM TnIH, Siemens Healthcare Diagnostics Inc., Tarrytown, NY, USA), run on an Atellica® IM Analyser. The LoB of the Atellica® IM TnIH assay was 0.50 ng/L. The observed LoD ranged from 1.13 to 1.53 ng/L across three reagent lots and two matrices (serum and lithium heparin plasma). The limit of quantitation (LoQ) of the Atellica® IM TnIH assay was 2.50 ng/l. The Atellica® IM TnIH assay provides results from 2.50 to 25,000.00 ng/L. The lower end of the measuring interval is defined by the LoQ. The Atellica® IM TnIH assay is a 3-site sandwich immunoassay using direct chemiluminometric technology. The solid phase reagent is based on magnetic latex particles conjugated with streptavidin with two bound biotinylated capture monoclonal antibodies, each recognising a unique cTnI epitope. The 99th percentile upper reference limit for healthy individuals was 34 ng/L for women and 53 ng/l for men. The total imprecision (CV) at the 99th percentile value of 45.20 pg/mL (ng/L) was below 10%.
Statistical methods

Standard descriptive statistics were applied in the analysis: the continuous variables were described as the mean ± SD and the median (5th percentile; 95th percentile), whereas the categorical variables were characterised by absolute and relative frequencies. The statistical significance of the differences among the groups of patients was analysed using the Mann-Whitney U test for continuous variables and the Fisher’s exact test for categorical variables. The contribution of the hs-cTnI biomarker to the NT-proBNP levels and clinical model was evaluated according to previously published recommendations [7, 8]. A logistic regression was adopted for the identification of predictors and development of multivariable models for scoring systems and biomarkers. The models were evaluated using a flexible calibration curve [9], C statistics, and a reclassification analysis of the model results. The analysis was completed using SPSS 24.0.0.1 (IBM Corporation, 2016) and R 3.5.1, with the PredictABEL and rms package.

Results

A total of 520 patients were evaluated. During the two-year follow-up, the primary endpoint occurred in 73 patients (14.0%). Of these patients, 55 died (without a previous LVAD implantation/HTX), 3 underwent LVAD implantations, 3 underwent LVAD implantations followed by HTX, 2 underwent LVAD implantation and later died, and 10 underwent HTX only. Acute heart failure hospitalization occurred in 74 patients (14.2%) of whom 30 patients (5.7%) experienced a further study endpoint and 44 (8.5%) heart failure hospitalization only. The secondary combined endpoint occurred in 117 patients (S1 Table). In general, the primary endpoint occurred more frequently in the patients who had a lower diastolic blood pressure (76 ± 9 vs. 81 ± 11 mmHg; p = 0.025), lower LVEF (27 ± 9 vs. 32 ± 9%; p < 0.001), and severe dyspnoea, i.e., those classified as NYHA III–IV (37.0% vs. 15.4%; p = 0.001). These patients also had higher hs-cTnI (34 vs. 17 ng/L; p < 0.001), NT-proBNP (1,950 vs. 518 ng/L; p < 0.001), and urea (8 vs. 6 mmol/L; p < 0.001) levels, and lower haemoglobin (135 vs. 145 g/L; p = 0.002) levels. The primary endpoint also occurred more frequently in the patients with a higher dose of furosemide (≥ 40 mg/day), but their heart failure treatment was otherwise comparable to the treatment given the patients without the primary endpoint (Table 1). A comparison of the characteristics of patients with/without the secondary endpoint is provided in S2 Table. Both groups differed in similar parameters as in the primary endpoint analysis (only statistically higher systolic BP was found in patients with the secondary endpoint, diastolic BP was comparable between groups). The patients with the secondary endpoint were also found to have significantly higher hs-cTnI levels (31 vs 16 ng/L; p < 0.001).

Relationship of hs-cTnI to clinical and laboratory parameters

The hs-cTnI levels differed significantly among the groups of patients categorized by age, LVEF, and renal function (expressed by the eGFR). In contrast, there were no significant differences found among the groups classified by sex, BMI, the ischaemic aetiology of the patients’ heart failure, hypertension, atrial fibrillation, diabetes mellitus, chronic obstructive pulmonary disease, or lower extremity peripheral artery disease (S3 Table). According to the Spearman’s rank correlation coefficient, the hs-cTnI levels were most strongly linked to the NT-proBNP levels (r = 0.439; p < 0.001) and renal function (the correlation coefficients for creatinine, urea, and eGFR were r = 0.245, 0.241, and -0.240, respectively, with p < 0.001 for all three parameters). The link was somewhat weaker when LVEF (r = -0.208), heart rate (r = 0.198), and age (r = 0.186) were considered (p < 0.001 for all three parameters) (S4 Table).
The prognostic value of hs-cTnI

Based on the ROC analysis (Table 2), we determined that 17 ng/L is the optimum cut-off value for the hs-cTnI level for predicting the monitored primary endpoint. The area under the curve (AUC) in this model was 0.658 (p < 0.001) (S1A Fig). The ROC analysis of hs-cTnI as a continuous parameter did not demonstrate any statistical significance. Fig 1 shows the distribution
of the hs-cTnI levels and the primary endpoint occurrence in the individual categories, according to the increasing levels of hs-cTnI. Patients with hs-cTnI levels < 17 ng/L are at a low risk of endpoint occurrence, whereas those with hs-cTnI levels ≥ 17 ng/L are at a high risk. It should be noted, however, that a further increase in troponin levels does not translate into a significant increase in the risk of primary endpoint occurrence. Patients with hs-cTnI levels < 17 ng/L accounted for 46.4% of all the patients. The monitored primary endpoint did not occur in any of the patients with hs-cTnI levels lower than the LoB (< 3 ng/L), but these patients only accounted for 2.9% of the total patients.

Similarly, based on the ROC analysis, we determined that 17 ng/L is also the optimum cut-off value for the hs-cTnI level for predicting the secondary endpoint. The area under the curve (AUC) in this model was 0.647 (p < 0.001) (S1B Fig).

Table 2. NT-proBNP [ng/l] and hs-cTnI [ng/l] as predictors of the primary endpoint in the logistic regression models (i.e., the two-year prognosis in terms of all-cause mortality, heart transplantation and left ventricular assist device [LVAD] implantation).

| Predictor                          | OR (95% CI)     | P    | AUC (95% CI) |
|-----------------------------------|-----------------|------|--------------|
| **Univariable models:**           |                 |      |              |
| NT-proBNP–categorical            | 1-category increase* | 1.92 (1.56; 2.36) | < 0.001 | 0.749 (0.695; 0.803) |
| hs-cTnI                           | 100-unit increase | 1.04 (0.98; 1.12) | 0.214  | 0.663 (0.601; 0.725) |
| hs-cTnI–binary                    | ≥ 17 (ref. < 17) | 4.35 (2.36; 8.01) | < 0.001 | 0.658 (0.596; 0.721) |
| **Multivariable model (combination of NT-proBNP and hs-cTnI):** |               |     |              |
| NT-proBNP–categorical            | 1-category increase* | 1.76 (1.42; 2.18) | < 0.001 | 0.767 (0.715; 0.819) |
| hs-cTnI [ng/l]–binary             | ≥ 17 (ref. < 17) | 2.30 (1.20; 4.41) | 0.012  | }
Multivariable prognostic models for primary endpoint

The AUC value for the prediction of the primary endpoint based on the ROC analysis of the NT-proBNP levels was 0.749 (p < 0.001) for the categorical model. The predictive value of the NT-proBNP plus hs-cTnI model was 0.767 (p < 0.001), according to the AUC (Table 2).

Based on previously published prognostic models, we developed a multivariable model by selecting comorbidities and clinical and laboratory parameters that have been shown to influence the prognosis of chronic heart failure patients [3, 4, 10]. Using a univariable logistic regression, we identified parameters that were significantly related to the monitored primary endpoint (Table 3). These parameters, together with the NT-proBNP and hs-cTnI levels, were used as inputs for the multivariable logistic regression. The significance of the individual variables selected for the multivariable model is shown in S1 Fig. The final multivariable model, which included NT-proBNP levels, NYHA (> 2), urea, and hs-cTnI levels ($\geq$ 17 ng/l), was created using a backward stepwise algorithm (Table 4). The excellent discriminatory power of the model to distinguish between patients with good and poor prognoses, assessed by the ROC analysis and expressed by the AUC, was 0.823 (p < 0.001) (S5 Table). Adding other parameters

Table 3. Patient characteristics as predictors of the primary endpoint in the univariable logistic regression models (i.e., the two-year prognosis in terms of all-cause mortality, heart transplantation and left ventricular assist device [LVAD] implantation).

| Predictor | OR (95% CI) | P |
|-----------|-------------|---|
| Sex | Men (ref. women) | 1.02 (0.54; 1.91) | 0.954 |
| Age | $\geq$ 65 (ref. < 65) | 2.01 (1.19; 3.41) | 0.009 |
| BMI | $\geq$ 30 (ref. < 30) | 1.15 (0.69; 1.91) | 0.585 |
| SBP [mmHg] | $\geq$ 110 (ref. < 110) | 0.40 (0.20; 0.81) | 0.010 |
| DBP [mmHg] | $\geq$ 80 (ref. < 80) | 0.54 (0.33; 0.89) | 0.015 |
| Heart rate [min$^{-1}$] | $\geq$ 75 (ref. < 75) | 1.95 (1.18; 3.23) | 0.009 |
| LVEF [%] | $\geq$ 35 (ref. < 35) | 0.36 (0.21; 0.65) | < 0.001 |
| Ischaemic aetiology of HF | Yes (ref. no) | 1.02 (0.62; 1.67) | 0.945 |
| Hypertension | Yes (ref. no) | 1.54 (0.88; 2.69) | 0.130 |
| Atrial fibrillation | Yes (ref. no) | 1.05 (0.62; 1.77) | 0.848 |
| Diabetes mellitus | Yes (ref. no) | 1.82 (1.11; 2.99) | 0.018 |
| COPD | Yes (ref. no) | 2.03 (1.12; 3.69) | 0.020 |
| Lower extremity peripheral artery disease | Yes (ref. no) | 1.91 (0.93; 3.93) | 0.079 |
| NYHA classification | $> 2$ (ref. $\leq 2$) | 3.78 (2.27; 6.29) | < 0.001 |
| Anaemia | Yes (ref. no) | 2.36 (1.35; 4.10) | 0.002 |
| Natrium [mmol/l] | $\geq$ 135 (ref. < 135) | 0.35 (0.14; 0.88) | 0.026 |
| Urea [mmol/l] | $\geq$ 6 (ref. < 6) | 4.19 (2.33; 7.52) | < 0.001 |
| Uric acid [μmol/l] | $\geq$ 500 (ref. < 500) | 2.82 (1.61; 4.96) | < 0.001 |
| Creatinine [μmol/l] | $\geq$ 100 (ref. < 100) | 2.09 (1.27; 3.44) | 0.004 |
| eGFR [ml/min/1.73 m$^2$] | $\geq$ 60 (ref. < 60) | 0.55 (0.33; 0.91) | 0.020 |
| ACEI/ARB | Yes (ref. no) | 0.56 (0.28; 1.11) | 0.096 |
| Beta-blockers | Yes (ref. no) | 0.63 (0.26; 1.50) | 0.297 |
| Furosemide $\geq$ 40 mg/day | $\geq$ 40 mg/day (ref. < 40) | 3.10 (1.73; 5.56) | < 0.001 |
| Spironolactone/eplerenone | Yes (ref. no) | 1.45 (0.84; 2.49) | 0.180 |

BMI, body mass index; COPD, chronic obstructive pulmonary disease; DBP, diastolic blood pressure; HF, heart failure; LVEF, left ventricular ejection fraction; SBP, systolic blood pressure; eGFR, estimated glomerular filtration rate (using the CKD-EPI equation).

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to this model did not improve its discriminatory power. After performing the 10-fold cross-validation to validate the results, the average AUC of 0.804 (in the range 0.689–0.892) was obtained. The calibration of the model was assessed by a flexible calibration curve, which confirmed that the predictive model is well calibrated (S3 Fig).

Our previously published multivariable model for predicting the monitored two-year end-point (overall mortality, HTX, or LVAD implantation) is based solely on clinical parameters and NT-proBNP levels [4]. This model includes the following independent parameters: older age, advanced heart failure (NYHA III+IV), anaemia, hyponatraemia, hyperuricaemia, and a higher dose of furosemide (> 40 mg daily). According to the ROC statistics and expressed by the AUC, the predictive power of this model, applied to the population of 520 patients included in this analysis, was 0.777 (0.722; 0.832) (p < 0.001).

A reclassification analysis confirmed the improvement of the predictive power of the model that combined the NT-proBNP and hs-cTnI levels, urea, and NYHA, as opposed to the model based solely on the NT-proBNP level and clinical parameters: the category-free net reclassification improvement was 0.473 (0.224; 0.723) (p < 0.001) and the integrated discrimination improvement was 0.054 (0.014; 0.094) (p = 0.008).

The final model was visualised using a nomogram (Fig 2), which makes it possible to establish the risk for each patient individually, based on their NT-proBNP and hs-cTnI levels, urea, and NYHA.

### Multivariable prognostic model for secondary endpoint

To create a multivariable model for secondary endpoint prediction, we proceeded similarly as in the primary endpoint analysis. Using a univariable logistic regression, we identified parameters that were significantly related to the occurrence of the secondary endpoint (age, systolic blood pressure, diastolic blood pressure, left ventricular ejection fraction, hypertension, diabetes mellitus, chronic obstructive pulmonary disease, NYHA classification, anaemia, natrium, urea, uric acid, estimated glomerular filtration rate, use of ACE inhibitors/angiotensin II receptor blockers, beta-blockers and higher dose of furosemide; the data are not presented). These parameters, together with NT-proBNP and hs-cTnI levels, were used as inputs for the multivariable logistic regression. The final multivariable model was created using a backward stepwise algorithm that included NYHA (> 2), NT-proBNP and urea levels. The level of hs-cTnI was not found to be an independent predictor of the secondary end-point (S6 Table). The discriminatory power of the model to distinguish between patients with good and poor prognosis, assessed by the ROC analysis and expressed by the AUC, was 0.802 (p < 0.001).

| Predictor | OR (95% CI) | P     |
|-----------|------------|-------|
| NYHA      |            |       |
| > 2 (ref. ≤ 2) | 3.24 (1.84; 5.68) | < 0.001 |
| hs-cTnI [ng/l] | ≥ 17 (ref. < 17) | 2.13 (1.08; 4.18) | 0.028 |
| NT-proBNP [ng/l] | 1-category increase* | 1.60 (1.29; 2.00) | < 0.001 |
| Urea [mmol/l] | 1-category increase** | 2.18 (1.47; 3.24) | < 0.001 |

*Categories of NT-proBNP: < 100 / 100–249 / 250–499 / 500–999 / 1,000–1,999 / ≥ 2,000.

**Categories of urea: < 6 / 6–9.9 / ≥ 10.

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Comparison of HFrEF and HFmrEF patient subpopulations

In our opinion, it is interesting to highlight some differences between the HFrEF and HFmrEF patient subpopulations. Men outnumber women in both groups; however, the proportion of women is higher in the HFmrEF group than in the HFrEF group (30% and 16.2%, respectively). Furthermore, HFmrEF patients are older, have a higher systolic blood pressure, and their heart failure is of ischaemic aetiology much more frequently. Overall, patients in the HFmrEF group are less frequently classified as NYHA III-IV, have lower median levels of hs-cTnl, NT-proBNP and uric acid, and are less frequently treated with furosemide $\geq 40$ mg/day and with mineralocorticoid receptor antagonists (S7 Table). According to Kaplan-Meier curves, the two-year survival without the monitored endpoint is 90.7 (85.1–96.3) in the HRmrEF group, and 82.7 (78.7–86.7) in the HRrEF group. According to the ROC analysis of our multivariable model, the area under the curve (AUC) is 0.860 (N = 120) for HRmrEF and 0.813 (N = 400) for HRrEF. The cut-off values for hs-TnI for HRrEF and HRmrEF patient subpopulations were determined to be 17 ng/l (N = 400) and ideally 14 ng/l (N = 120), respectively; but even 17 ng/l is a very good cut-off value for the HRmrEF group. In view of the fact that the HRmrEF group involves a markedly lower number of patients, we consider a single cut-off value to be adequate, namely 17 ng/l.

Discussion

This is the first study to evaluate the prognostic significance of cardiac troponin I (Atellica IM High sensitivity Troponin I) in patients with chronic heart failure. This paper presents
three important results. First, the hs-cTnI cut-off value for predicting a poorer prognosis of HFrEF/HFmrEF patients, as determined by Atellica®, is 17 ng/L, and a further increase in the troponin level is not linked to a significant risk increase. Second, the hs-cTnI levels in chronic heart failure patients are linked to their age, LVEF, and renal function, and they correlate with the NT-proBNP levels. Third, although the predictive value of the hs-cTnI level by itself, as expressed by the C-statistics, is relatively low, adding the hs-cTnI level to the NT-proBNP level and clinical parameters significantly improves the identification of patients that are at a higher risk of a combined endpoint (death, heart transplantation or LVAD implantation). From both analytical and clinical points of view, the value of hs-cTnI under 17 ng/l also helps to identify low-risk patients. Nevertheless, the value of hs-cTnI does not increase the power of the model for secondary endpoint prediction that includes decompensated heart failure hospitalizations (on top of primary endpoint components).

**Clinical benefit of the model and its interpretation**

Our model makes it possible to identify the highest-risk patients in the population of HFrEF/HFmrEF patients. This subgroup of patients should be followed up in specialised centres for chronic heart failure patients and, in accordance with current guidelines, re-examinations, and possibly intensification of pharmacological treatments (increasing the dose of diuretics/ACE inhibitors/sartans/beta-blockers; replacing ACE inhibitors/sartans with sacubitril/valsartan; adding spironolactone/epi-lerone, ivabradine, SGLT2 inhibitors or digoxin) and/or non-pharmacological treatments (revascularisation, ICD/CRT implantation, valvular heart disease surgery), might be considered. As a last resort, LVAD implantation or putting the patient on a waiting list for heart transplantation may also need to be considered [2].

From a practical standpoint, it is reasonable to find the boundary beyond which patients are at risk. During the two-year follow-up, the primary endpoint occurred in 14% of the patients in our cohort. In the group of patients with advanced heart failure, for whom the benefit of LVAD implantation was not yet definite (INTERMACS profile 6–7), the two-year occurrence of death/orthotopic heart transplantation/LVAD implantation was 42%. In the group of patients with an INTERMACS profile of 4–5, for whom the benefit of LVAD implantation (as compared with the optimum pharmacotherapy) had been demonstrated, the two-year occurrence of the combined endpoint was 53% [11]. In contrast, the two-year overall mortality rate of patients with advanced heart failure who had undergone the implantation of a centrifugal-flow pump (HeartMate 3) was 11.9%, and the survival at two years, free of disabling stroke or reoperation to replace or remove a malfunctioning device, was 79.5% [12].

With the use of the suggested predictive model, patients can be divided into four groups: low-risk patients (up to 10%, with a score of up to 120 points), low-to-moderate-risk patients (11%–30%, with a score between 121 and 180 points), moderate-to-high risk (30%–45%, with a score between 181 and 208 points), and high-risk patients (45%, with a score of 209 or more points). Patients with extremely low hs-cTnI levels (< 3 ng/L) have a very good prognosis.

Intensification of pharmacological/non-pharmacological treatments might be considered for the group of low-to-moderate-risk patients, whereas moderate-to-high-risk and high-risk patients should be followed up in centres specialised in heart failure.

**Comparison with other models**

The main advantages of our model are as follows: (1) the use of a combined endpoint, where LVAD implantation and heart transplantation are monitored together with mortality and (2) the use of routinely available biomarkers. We have demonstrated the benefit of adding the hs-cTnI level to the NT-proBNP level and clinical parameters. The two-year mortality prediction...
was published in the Seattle Heart Failure Model, with an AUC of 0.792 [13]. The NT-proBNP level was also used in a recently published model for a two-year prediction of a primary composite endpoint (cardiovascular death or hospitalisation for heart failure) (AUC 0.71), cardiovascular death (AUC 0.71), and all-cause death (0.70) [5]. The use of biomarkers, particularly natriuretic peptides, should be a standard procedure to establish the degree of risk in heart failure patients. The NT-proBNP level was the most significant prognostic factor, not only in our study, but also in an extensive cohort of patients in the PARADIGM-HF study [5]. Our results are in accordance with a recently published, and rather extensive, meta-analysis that evaluated the prognostic value of hs-cTnT for individual patients with chronic heart failure. An hs-cTnT level \( \geq 18 \) ng/L, added to clinical parameters and the NT-proBNP level, was an independent predictor for a more than two-year prognosis of monitored endpoints (all-cause mortality, cardiovascular mortality, and cardiovascular hospitalisation) [6].

Our study showed that including an acute heart failure hospitalization as the component of the combined endpoint leads to a diminished predictive power of increased hs-cTnI value. Similar findings were demonstrated by previous analyses of risk models in patients with heart failure published by Rahimi et al. The discriminatory ability of the models for prediction of death appeared to be higher than that for prediction of death or acute heart failure hospitalisation or prediction of hospitalizations alone [3].

Limitations

Our study has several limitations. First, we only described a population of patients with mid-range and reduced ejection fractions; any patient with a preserved ejection fraction was excluded from the study. Due to the limited number of patients, we did not attempt to establish whether men and women might have different cut-off values for predicting a poor prognosis. Our model was not validated in an external cohort of patients. Additionally, our results respond to hs-cTnI measurements in chronic heart failure patients in a stable phase in outpatient clinics. And finally, population sampling might be influenced by a selection bias. This group of heart failure patients has been followed up in university centres; therefore, the patient cohort very probably was not a representative sample of heart failure patients with EF \(<50\%\) from across the Czech Republic. A study from the Netherlands, based on an extensive registry of heart failure patients, which was intended to describe the real-world population with heart failure [14], revealed that about 77% of all patients had EF \(<50\%\). Their mean age was 72 years (as compared to 65 years in our cohort) and men accounted for 64% only (as compared to 80.6% in our cohort).

Conclusion

Hs-cTnI levels \( \geq 17 \) ng/l represent the cut-off for predicting a poorer prognosis for HFrEF/HFmrEF patients. We have developed a simple model for the early identification of HFrEF/HFmrEF patients whose prognoses might be improved if an intensification of treatment is considered (based on the indication criteria). This treatment intensification could include putting the patients on a waiting list for heart transplantation or LVAD implantation.

Supporting information

S1 Fig. A: ROC curve for prediction of combined primary endpoint (death, HTX or LVAD) using hsTnI; B: ROC curve for prediction of combined secondary endpoint (death, HTX, LVAD and/or acute heart failure hospitalization) using hsTnI.

(TIF)
S2 Fig. Significance of the variables as measured using a partial Wald $\chi^2$ test minus the predictor degrees of freedom in the full model and the final model selected by a backward stepwise algorithm.

(TIF)

S3 Fig. Calibration curve of the multivariable logistic regression model.

(TIF)

S1 Table. Summary of event occurrence in the first two years of follow-up.

(DOCX)

S2 Table. Basic characteristics of the patients according to occurrence of secondary endpoint (i.e., the two-year prognosis in terms of all-cause mortality, heart transplantation, left ventricular assist device [LVAD] implantation, hospitalization for HF).

(DOCX)

S3 Table. Comparison of hs-cTnI between patient subgroups.

(DOCX)

S4 Table. Correlation between hs-cTnI and other parameters.

(DOCX)

S5 Table. Accuracy of the multivariable prediction model.

(DOCX)

S6 Table. The multivariable logistic regression model using a backward stepwise algorithm for the selection of independent predictors of the primary endpoint (i.e., the two-year prognosis in terms of all-cause mortality, heart transplantation, left ventricular assist device [LVAD] implantation, hospitalization for HF).

(DOCX)

S7 Table. Basic characteristics of patients with HFrEF/HFmrEF.

(DOCX)

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References

1. Parenica J, Spinar J, Vitovec J, Widimsky P, Linhart A, Fedorco M, et al. Long-term survival following acute heart failure: the Acute Heart Failure Database Main registry (AHEAD Main). Eur J Intern Med. 2013; 24: 151–160. https://doi.org/10.1016/j.ejim.2012.11.005 PMID: 23219321

2. Ponikowski P, Voors AA, Anker SD, Bueno H, Cleland JGF, Coats AJS, et al. 2016 ESC Guidelines for the diagnosis and treatment of acute and chronic heart failure: The Task Force for the diagnosis and treatment of acute and chronic heart failure of the European Society of Cardiology (ESC) Developed with the special contribution of the Heart Failure Association (HFA) of the ESC. Eur Heart J. 2016; 37: 2129–2200. https://doi.org/10.1093/eurheartj/ehw128 PMID: 27206819

3. Rahimi K, Bennett D, Conrad N, Williams TM, Basu J, Dwight J, et al. Risk prediction in patients with heart failure: a systematic review and analysis. JACC Heart Fail. 2014; 2: 440–446. https://doi.org/10.1016/j.jchf.2014.04.008 PMID: 25194291

4. Spinar J, Spinarova L, Malek F, Ludka O, Krejci J, Ostadal P, et al. Prognostic value of NT-proBNP added to clinical parameters to predict two-year prognosis of chronic heart failure patients with mid-range and reduced ejection fraction—A report from FAR NHL prospective registry. PloS One. 2019; 14: e0214363. https://doi.org/10.1371/journal.pone.0214363 PMID: 30913251

5. Simpson J, Jhund PS, Lund LH, Padmanabhan S, Claggett BL, Shen L, et al. Prognostic Models Derived in PARADIGM-HF and Validated in ATMOSPHERE and the Swedish Heart Failure Registry to Predict Mortality and Morbidity in Chronic Heart Failure. JAMA Cardiol. 2020. https://doi.org/10.1001/jamacardio.2019.5850 PMID: 31995119

6. Aimo A, Januzzi JL, Vergaro G, Ripoli A, Latini R, Masson S, et al. Prognostic Value of High-Sensitivity Troponin T in Chronic Heart Failure: An Individual Patient Data Meta-Analysis. Circulation. 2018; 137: 286–297. https://doi.org/10.1161/CIRCULATIONAHA.117.031560 PMID: 29392588

7. Hlatky MA, Greenland P, Arnett DK, Ballantyne CM, Criqui MH, Eikind MSV, et al. Criteria for evaluation of novel markers of cardiovascular risk: a scientific statement from the American Heart Association. Circulation. 2009; 119: 2408–2416. https://doi.org/10.1161/CIRCULATIONAHA.109.192278 PMID: 19364974

8. Wang TJ. Assessing the role of circulating, genetic, and imaging biomarkers in cardiovascular risk prediction. Circulation. 2011; 123: 551–565. https://doi.org/10.1161/CIRCULATIONAHA.109.912568 PMID: 21300963

9. Van Calster B, Nieboer D, Vergouwe Y, De Cock B, Pencina MJ, Steyerberg EW. A calibration hierarchy for risk models was defined: from utopia to empirical data. J Clin Epidemiol. 2016; 74: 167–176. https://doi.org/10.1016/j.jclinepi.2015.12.005 PMID: 26772608

10. Spinar J, Jarkovsky J, Spinarova L, Mebazaa A, Gayat E, Vitovec J, et al. AHEAD score—Long-term risk classification in acute heart failure. Int J Cardiol. 2016; 202: 21–26. https://doi.org/10.1016/j.ijcard.2015.08.187 PMID: 25386914

11. Ambardekar AV, Kittleson MM, Palardy M, Mountis MM, Forde-McLean RC, DeVore AD, et al. Outcomes with ambulatory advanced heart failure from the Medical Arm of Mechanically Assisted Circulatory Support (MedaMACS) Registry. J Heart Lung Transplant. 2019; 38: 408–417. https://doi.org/10.1016/j.healun.2018.09.021 PMID: 30948210
12. Mehra MR, Goldstein DJ, Uriel N, Joseph C, Cleveland J, Yuzefpolskaya M, Salerno C, et al. Two-Year Outcomes with a Magnetically Levitated Cardiac Pump in Heart Failure. N Engl J Med. 2018 [cited 1 Mar 2020]. https://doi.org/10.1056/NEJMoa1800866 PMID: 29526139

13. Levy WC, Mozaffarian D, Linker DT, Sutradhar SC, Anker SD, Cropp AB, et al. The Seattle Heart Failure Model: prediction of survival in heart failure. Circulation. 2006; 113: 1424–1433. https://doi.org/10.1161/CIRCULATIONAHA.105.584102 PMID: 16534009

14. Brunner-La Rocca H-P, Linssen GC, Smeele FJ, van Drimmen AA, Schaafsma H-J, Westendorp PH, et al. Contemporary Drug Treatment of Chronic Heart Failure With Reduced Ejection Fraction: The CHECK-HF Registry. JACC Heart Fail. 2019; 7: 13–21. https://doi.org/10.1016/j.jchf.2018.10.010 PMID: 30606482