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Gene dysregulation analysis builds a mechanistic signature for prognosis and therapeutic benefit in colorectal cancer

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The implementation of cancer precision medicine requires biomarkers or signatures for predicting prognosis and therapeutic benefits. Most of current efforts in this field are paying much more attention to predictive accuracy than to molecular mechanistic interpretability. Mechanism-driven strategy has recently emerged, aiming to build signatures with both predictive power and explanatory power. Driven by this strategy, we developed a robust gene dysregulation analysis framework with machine learning algorithms, which is capable of exploring gene dysregulations underlying carcinogenesis from high-dimensional data with cooperativity and synergy between regulators and several other transcriptional regulation rules taken into consideration. We then applied the framework to a colorectal cancer (CRC) cohort from The Cancer Genome Atlas. The identified CRC-related dysregulations significantly covered known carcinogenic processes and exhibited good prognostic effect. By choosing dysregulations with greedy strategy, we built a four-dysregulation (4-DysReg) signature, which has the capability of predicting prognosis and adjuvant chemotherapy benefit. 4-DysReg has the potential to explain carcinogenesis in terms of dysfunctional transcriptional regulation. These results demonstrate that our gene dysregulation analysis framework could be used to develop predictive signature with mechanistic interpretability for cancer precision medicine, and furthermore, elucidate the mechanisms of carcinogenesis.

Keywords: gene dysregulation analysis, mechanistic signature, cancer precision medicine, prognosis, chemotherapy benefit, colorectal cancer

Introduction

Biomarkers or signatures for predicting prognosis and therapeutic benefits are an indispensable part for implementing cancer precision medicine (Walther et al., 2009; Vargas and Harris, 2016). The improvements of prognostic and therapeutic benefits with the aid of signatures have been reported in colorectal cancer (CRC), breast cancer, lung cancer, etc. (Chen et al., 2007; Khambata-Ford et al., 2007; Salazar et al., 2011; Sparano et al., 2019). Meta-analysis studies of clinical trials demonstrate that response rate seen with targeted agents under biomarker guidance has reached \( \sim 30\% \), which is much higher than that of chemotherapies (Schwaederle et al., 2015). However, despite these promising outcomes, there remain urgent needs for improving the performance of clinical signatures.

The most widely adopted strategies for identifying predictive signatures heavily depend on expression analysis of individual genes involving identification of differentially expressed genes (DEGs) (Khambata-Ford et al., 2007; Salazar et al., 2011) and genes whose expression values are relevant to certain
phenotype (Chen et al., 2007; Sparano et al., 2019). It is apparent that they ignore gene interconnection implicated in transcriptomic data, even though genes perform their functions in coordination, instead of in isolation (Barzel and Barabasi, 2013). Therefore, the currently available predictive signatures unavoidably tend to have limited mechanism explanatory power, which has become a common concern in precision medicine (Robinson et al., 2013). There is a general consensus among both clinicians and biologists about the need for signatures with mechanistic interpretability as well as high predictive accuracy for cancer precision medicine (Robinson et al., 2013; Topalian et al., 2016; Lu et al., 2019). It could also be expected that taking mechanistic interpretation into consideration would further enhance the predictive accuracy and robustness of signatures in clinical application (Robinson et al., 2013).

Attributed to the crucial roles of gene regulation in fundamental cell processes, to build signatures with mechanistic interpretability requires studying relevant genes in the context of regulatory networks, that is, identifying specific regulators and their regulatory relationships that are dysfunctional in a given disease state (Lee and Young, 2013). In the past >10 years, quite a few investigations have aimed to elucidate dysfunctional regulatory networks in disease, instead of solely focusing on DEGs (de la Fuente, 2010; Ideker and Krogan, 2012). Among them, differential coexpression analysis (DCA, also shortened as ‘DCEA’ in literatures), which were developed to identify differences in genes coexpression patterns between healthy and disease samples, was regarded as the first steps toward differential regulation analysis or gene dysregulation analysis (de la Fuente, 2010). Most of the existing DCA-based methodologies construct coexpression networks and identify the alteration of gene–gene identities (i.e. network topology) or gene–gene correlation (i.e. edge weight) (Ideker and Krogan, 2012; Li et al., 2016). Based on these strategies, differential modules or gene sets related to certain phenotypes could be identified to build signatures for predicting prognosis (Taylor et al., 2009) and drug response (Zickenrott et al., 2016). However, most signatures involve too many genes and include too much noise, which greatly weaken their mechanism explanatory power. In our early work, we applied our previously developed DCA-based method to cancer and generated a series of carcinogenesis-relevant biomarkers (Liu et al., 2010; Yu et al., 2011; Yang et al., 2013; Wu et al., 2016; Li et al., 2017a; Dai et al., 2018). Benefiting from the quantitative design of our DCA-based algorithms and candidate gene screening by transcriptional regulation relationships, the numbers of genes in our biomarkers were decreased to a practical level and the mechanistic interpretability of the identified biomarkers were also elevated (Wu et al., 2016; Li et al., 2017a).

Since correlation analysis cannot distinguish direct associations from indirect associations (Barzel and Barabasi, 2013), it has limited potential to directly provide clues into disease mechanisms. Aiming to identify the alterations of gene regulation relationships instead of expression correlations between various phenotypes, updated versions of differential regulation analysis have emerged as a follow-up effort toward elucidating disease-related dysfunctional regulatory networks or differential regulation relationships (Cao et al., 2015; Li et al., 2017b). Considering that transcriptional regulation requires sufficient cooperativity and synergy of multiple regulators (Lambert et al., 2018), however, it is still a huge challenge about how to robustly build a framework for dysregulation analysis based on high-dimensional transcriptome data if taking cooperativity and synergy of multiple regulators into consideration.

In this work, we first proposed a framework for gene dysregulation analysis by using machine learning algorithms that are able to consider the cooperativity and synergy between regulators and robustly cope with high-dimensional data (Figure 1). A reference gene regulation network (GRN) was constructed by predicting the potential binding site of transcription factors (TFs) among promoter regions (Figure 1A). Conditional GRNs were then highlighted with a random forest-based feature selection algorithm Boruta (Kursa and Rudnicki, 2010; Figure 1B), and each link’s regulatory intensity and its confidential interval were estimated with a de-biased least absolute shrinkage and selection operator (LASSO) method (Javanmard and Montanari, 2010). Conditional GRNs could then be validated of the quantifying method of regulatory intensity).
Subsequently, 389 gene dysregulations were extracted according to three factors including differential regulation, differential expression of target, and the consistency between differential regulation and differential expression (Supplementary Table S1). Two examples of gene dysregulation, RUNX3—GPR15 and KLF6—WNT2, were illustrated in Figure 2. For RUNX3—GPR15 (Figure 2A), the regulatory intensity was reduced from normal to cancer with the intensity value being 0.441 (normal) and –0.010 (cancer), and the expression level of the target, GPR15, significantly decreased as shown. For KLF6—WNT2 (Figure 2B), the regulatory intensity was increased from normal (–0.383) to cancer (0.274), and accordingly, the expression level of the target, WNT2, was elevated. It was interesting that either the target set (341 genes) or the TF set (262 genes) involved in dysregulations was able to correctly classify tumor and normal samples with unsupervised hierarchical clustering method, suggesting that the 389 gene dysregulations were potentially relevant to carcinogenesis (Figure 2C; Supplementary Figure S1).

To globally understand the functions of the 389 dysregulations, we implemented pathway over-representation analysis and obtained numerous cancer-related pathways enriched for the dysregulation genes (Figure 2D; Supplementary Table S2). In order to explore the influences of dysregulation events on carcinogenesis, we extracted the targets’ expression data and checked the change of pathway activities with gene set enrichment analysis (GSEA) (Subramanian et al., 2005). It was shown that several biological processes including cell proliferation, cell cycle, pathways in cancer, chromosome organization, and vascular development were more active in CRC samples (Figure 2E for cell proliferation; Supplementary Figure S2 for all), while some others including immune system process, cell death, and cell–cell adhesion were inhibited in cancer (Figure 2F for immune system process; Supplementary Figure S2 for all). These results support that the identified 389 dysregulations essentially highlight cancer-related crucial processes and thus could be taken as functional seeds for building explanatory signature.

Prognostic effects of gene dysregulations

We first checked prognostic effects of the 389 gene dysregulations for overall survival/recurrent-free survival (OS/RFS) on TCGA CRC dataset. Following the procedures in Materials and methods, we found that for all four types of cox models, C-indexes of 389 models fitted with the 389 dysregulations were significantly larger than those with 389 gene pairs randomly selected through four gene selection strategies (see Materials and methods for details), with the median P-value of 100 times of Wilcox tests <0.05 (Figure 3A). We then tested the prognostic effect of the 389 dysregulations in two independent datasets, GSE39582 and GSE17538. In GSE39582, similar trends were obtained excepting the OS cox model fitted with expression data and clinical information when the control setting involved DEGs (Figure 3B, comparison_2 and comparison_4 in Exp+Clin_OS). In GSE17538,
similar trends were also observed apart from the RFS cox model fitted with expression data alone (Figure 3C, Exp_RFS). These results indicate that the 389 CRC-related dysregulations have prognostic capability as a whole.

Construction of prognostic signature with dysregulations

Considering that our identified 389 dysregulations could cover known carcinogenic processes (Figure 2) and exhibit prognostic effect (Figure 3), we set out to construct a high-accuracy prognostic model based on these dysregulations, which was expected to provide both mechanism explanatory power and predictive power. First of all, among the 389 OS cox models fitted with expression data of 389 dysregulations adjuvant with clinical information from TCGA CRC dataset (Supplementary Table S3), the model for RUNX3—GPR15 dysregulation stood out with the largest C-index, 0.763. RUNX3, a member of the runt domain-containing family of TFs, has been found to be essential for diverse processes including proliferation, differentiation, cell lineage specification, apoptosis, and DNA repair (Ito et al., 2015; Bae et al., 2019). RUNX3 is taken
as a tumor suppressor in CRC (Weisenberger et al., 2006; Soong et al., 2009). GPR15, an orphan G-protein-linked receptor, was reported to mediate T-cells localization to colon (Habtezion et al., 2016), and very recently proposed to represent a therapeutic target for CRC (Namkoong et al., 2019).

Given that plasma membrane proteins show great potential as drug targets and diagnostic objectives, we narrowed down the 389 dysregulations to 126 plasma membrane protein-relevant dysregulations to build a prognostic signature with explanatory capability (Supplementary Table S3). Starting from RUNX3→GPR15, a greedy strategy as described in Materials and methods was used to select other dysregulations to boost the signature’s performance. At last, a prognostic signature involving four dysregulations (4-DysReg), RUNX3→GPR15, RUNX3→P2RY8, SNAI3→TLR7, and ATOH1→SIGLEC1, was built.

Based on the expression data of the seven genes involved in 4-DysReg (Supplementary Figure S3) and the clinical information including age, gender, and pathological stage in TCGA CRC dataset, we built an OS cox model on the entire samples from TCGA CRC dataset. C-index of this model ran up to 0.79 (SE = 0.038). The risk scores for every sample were then calculated. The OS time, survival status, and risk score were shown in Figure 4A, indicating a high correlation between risk score and survival status. Time-dependent receiver operating characteristic (ROC) curves showed that this model represents high accuracy for OS prediction, with the area under ROC curve (AUC) at 1-, 3-, and 5-year survival reaching 0.82, 0.79, and 0.78 (Figure 4B). After that, median risk score was used to cut the samples into high and low score groups, and patients with low score displayed significantly better prognosis (hazard ratio (HR) = 0.134; 95% CI: 0.071–0.253; P-value = 6.52e–10; Figure 4C). Following the same analysis procedure as in TCGA CRC dataset, the OS predictive power of 4-DysReg was also validated on independent datasets, GSE39582 and GSE17538 (Supplementary Figure S4). Besides, we checked the performance of 4-DysReg in various clinical stratification subtypes with regard to age, gender, primary site, pathological stage, lymphatic invasion, and microsatellite status. The results showed that low score group always had significantly longer survival time than high score group in almost every subtype (Figure 4D). It is noticed that patients in low score group showed significantly better prognosis in both microsatellite instability (MSI) subgroup (n = 107) and microsatellite stability (MSS) subgroup (n = 235) (Figure 4D). Since MSI is a mature biomarker in CRC for prognosis and for some chemotherapies (Boland and Goel, 2010), it seems that 4-DysReg has a good consistency across various subtypes for CRC prognosis prediction.

At last, cross-validation on three datasets, TCGA CRC, GSE39582, and GSE17538, indicated that C-index, 1-, 3-, and 5-year survival AUC, and log-rank test of the testing sets maintained high levels compared to the training sets (Table 1). Taken together, our signature 4-DysReg is capable of robustly predicting OS with high accuracy.

**Comparison of the predictive accuracy of 4-DysReg with other CRC signatures**

We compared the predictive accuracy of 4-DysReg for OS with previously reported CRC expression signatures, including RUNX3 (Soong et al., 2009), ColoPrint (contains 18 genes) (Salazar et al., 2011), ColoGuideEx (contains 13 genes) (Agesen et al., 2012), ColoGuidePro (contains 7 genes) (Sveen et al., 2012), ColoFinder (contains 9 genes) (Shi and He, 2016), CRCassigner-30 (Sadanandam et al., 2013),
Figure 4 Construction of prognostic signature base on dysregulations with TCGA CRC dataset. (A) Scatter plot for OS time, survival status, and risk score for TCGA CRC samples. (B) Time-dependent ROCs for risk score at 1-, 3-, and 5-year survival. (C) Kaplan–Meier curves of OS between two groups cut by median risk core. P-value was generated from log-rank test. (D) Forest plots of the associations between 4-DysReg risk score and OS in various subgroups. High score: the group with risk score larger than the median; low score: the group with risk score smaller than the median.
Table 1 Cross-validation of 4-DysReg in terms of C-indexes, AUC of time-dependent ROC at 1-, 3-, and 5-year survival, and log-rank test in three cohorts.

| Cohort          | C-index   | AUC_1   | AUC_3   | AUC_5   | −log10 (log-rank test P-value) |
|-----------------|-----------|---------|---------|---------|-------------------------------|
| TCGA CRC        | Training  | 0.80 (0.78–0.81) | 0.83 (0.80–0.85) | 0.80 (0.78–0.83) | 0.79 (0.76–0.82) | 7.74 (6.23–8.81) |
|                 | Testing   | 0.76 (0.73–0.78) | 0.79 (0.75–0.82) | 0.76 (0.72–0.80) | 0.73 (0.69–0.78) | 4.15 (3.23–5.01) |
| GSE39582        | Training  | 0.71 (0.69–0.72) | 0.70 (0.65–0.77) | 0.73 (0.69–0.82) | 0.75 (0.71–0.86) | 6.97 (5.85–8.12) |
|                 | Testing   | 0.67 (0.65–0.71) | 0.76 (0.72–0.79) | 0.71 (0.69–0.74) | 0.70 (0.68–0.72) | 3.41 (3.31–4.08) |
| GSE17538        | Training  | 0.75 (0.73–0.77) | 0.84 (0.80–0.87) | 0.78 (0.75–0.81) | 0.78 (0.76–0.81) | 3.52 (2.99–4.23) |
|                 | Testing   | 0.71 (0.68–0.75) | 0.76 (0.70–0.81) | 0.69 (0.64–0.73) | 0.71 (0.65–0.75) | 1.54 (1.00–2.08) |

Data are expressed as median (first quantile–third quantile) of 100 times of cross-validation.

Figure 5 Comparison of the predictive accuracy of 4-DysReg with other CRC expression signatures. (A) Results in TCGA CRC dataset. (B) Results in TCGA CRC dataset with clinical information included. (C) Results in GSE39582 dataset. (D) Results in GSE39582 dataset with clinical information included. In this cross-validation, 60% of the samples were randomly selected as training set to fit a cox model with each signature, and the left 40% were taken as testing set to calculate C-index. Cross-validation of each signature was repeated 100 times.

CRCassigner-7 (Sadanandam et al., 2013), a 7-gene signature (Chen et al., 2017), and a 4-gene signature (Zou et al., 2015). The detailed gene information of every signature could be found in Supplementary File. In TCGA CRC dataset, our 4-DysReg signature outperformed all the others in the cross-validation (Figure 5A). With clinical information included, the accuracy of all signatures was enhanced and 4-DysReg still performed the best (Figure 5B). In GSE39582 dataset, ColoGuideEx alone displayed a slightly higher C-index than 4-DysReg (Figure 5C), while 4-DysReg surpassed ColoGuideEx when combining clinical information (Figure 5D). Taken together, these observations indicate the predictive power of 4-DysReg for prognosis.

Predictive power of 4-DysReg for chemotherapeutic benefit

ADJC is preferred for curing CRC patients, and the guideline of ADJC is established on pathologic stage (Watanabe et al., 2018). Stage III and IV CRC patients are routinely recommended
to receive ADJC. Stage II CRC patients with high risk of recurrence also consider ADJC, but the usefulness of postoperative ADJC has not been proved (Watanabe et al., 2018).

Herein, we adopted a large-scale CRC dataset GSE39582 (553 samples), with ADJC records ranging from stage II to stage IV, to explore the predictive power of 4-DysReg for chemotherapeutic benefit regardless of pathologic stage. An OS cox model was trained with the expression data of the seven genes involved in 4-DysReg on samples without ADJC ( \( n = 321 \) ), which was used to calculate risk scores of samples with ADJC ( \( n = 232 \) ) (see Supplementary File for the rationality of choosing the training set). It was shown that the sample group with negative risk score had better prognosis (HR = 0.432; 95% CI: 0.269–0.693; Figure 6A).

Furthermore, we analyzed the predictive capability of 4-DysReg for benefit of specific chemotherapy types including 5-FU and combined ADJC (including FOLFI, FOLFOX, and FUFA). Still, the samples with negative risk score basically showed better survival for both combined ADJC ( \( n = 84 \); HR = 0.380; 95% CI: 0.178–0.813; Figure 6B) and 5-FU ( \( n = 79 \); HR = 0.437; 95% CI: 0.181–1.055; Figure 6C). We also compared the predictive capability of 4-DysReg for chemotherapeutic benefit with previously reported signatures. Excepting CRCassigner7 for 5-FU chemotherapy benefit, 4-DysReg performed much better when predicting chemotherapeutic benefit of all ADJC, combined ADJC, and 5-FU (Supplementary File). The accuracy of the signature for combined ADJC was evaluated by time-dependent ROC, and 3- and 5-year survival AUC reached 0.71 and 0.74, higher than AUC of model based on pathologic stage, 0.65 and 0.69 (Figure 6D).

These results demonstrate that our 4-DysReg signature possesses the predictive power for therapeutic benefit of ADJC, including only for 5-FU, only for combined ADJC, and for all ADJC, and therefore has the potential to guide ADJC regardless of pathologic stage information.

Discussion

In the field of cancer precision medicine, a series of signatures have been built for prognosis or therapeutic benefits based on gene expression data (Chen et al., 2007; Khambata-Ford et al., 2007; Salazar et al., 2011; Sparano et al., 2019). However, current efforts are always focusing on predictive accuracy over explanatory power. Driven by the recent mechanism-driven strategies, we developed a robust gene dysregulation analysis framework by using machine learning algorithms, which is capable of robustly exploring gene dysregulations underlying carcinogenesis from high-dimensional data without ignoring cooperativity and synergy between regulators. We applied our framework on a TCGA CRC cohort and eventually developed a dysregulation-based signature with the capability of predicting prognosis and chemotherapy benefit as well. Our signature 4-DysReg showed superior performance over a series of previously reported CRC expression signatures for prognosis and chemotherapy benefit. Compared with those reported signatures generated from individual gene expression analysis based methodology (Xiong
et al., 2018), 4-DysReg involves much fewer genes, and have more functional relevance in carcinogenesis in terms of dysfunctional transcriptional regulation.

It has been widely accepted that dysfunctional regulation of gene expression programs can cause a broad range of phenotypic changes, such as carcinogenesis (Lee and Young, 2013). Systematic identification of gene dysregulations based on omics data is an effective path for exploring molecular mechanisms behind phenotypic changes. Quite a few approaches have been reported to address this by measuring the change of expression correlation between gene pairs across phenotypes (de la Fuente, 2010; Ideker and Krogan, 2012; Li et al., 2016), yet in the framework of correlation analysis, this pair-based estimation inevitably overrates the contribution of an individual regulator to its target. In order to simultaneously consider the contributions of multiple regulators to transcriptional regulations, even the cooperative and synergistic effect between multiple regulators (Lambert et al., 2018), we have previously developed dysregulation analysis methods based on multivariate regression, which were applied to cancer and generated carcinogenesis-relevant biomarkers (Cao et al., 2015; Li et al., 2017b). However, the methodology is far from robust when dealing with high-dimensional transcriptome data for dysregulation analysis on a systematic level.

In the present dysregulation analysis framework, the cooperative and synergistic effect between regulators could be fully considered in a robust way. First, conditional GRNs are constructed by selecting important TFs for target expression with feature selection algorithm Boruta, in which the relative importance of all candidate TFs are estimated simultaneously (Kursa and Rudnicki, 2010); secondly, regulatory intensities and its Cs of remaining TFs for target expression are quantified with de-biased LASSO (Javanmard and Montanari, 2014), where the expression of the target is determined by the combined regulatory effect of all possible regulators, which naturally considers the cooperativity and synergy between TFs. Besides, we identify gene dysregulations according to three factors including differential regulation, differential expression of target, and the consistency between differential regulation and differential expression. In this way, our dysregulation analysis framework is robust in dealing with high-dimensional transcriptome data and endowed with explanatory power by combining biological principles and machine learning algorithms with interpretability.

The explanatory signature 4-DysReg built in this study consists of four dysregulations including RUNX3→GPR15, RUNX3→P2RY8, SNAI3→TLR7, and ATOH1→SIGLEC1. The 4-DysReg showed good performance in predicting prognosis and therapeutic benefit of ADJIC in CRC patients. Attributed to dysregulation analysis, 4-DysReg has sufficient interpretability in terms of dysfunctional transcriptional regulation and molecular functions. As expected, the seven genes involved in 4-DysReg have been reported to participate in diverse processes related to carcinogenesis. RUNX3 is a member of the runt domain-containing family of TFs, and participates in diverse processes including proliferation, differentiation, cell lineage specification, apoptosis, and DNA repair (Ito et al., 2015; Bae et al., 2019). In CRC, RUNX3 is taken as a tumor suppressor since its methylation is a significant risk factor for tumor development and high nuclear expression of RUNX3 is associated with better survival (Weisenberger et al., 2006; Soong et al., 2009). It is intriguing that RUNX3 demonstrates both tumor-suppressive and oncogenic activities in multiple solid tumors (Ito et al., 2015). Still intriguingly, GPR15, an orphan G-protein-linked receptor, is recently reported to mediate T-cells localization to colon when expressed on cell surface of T cells (Habtezion et al., 2016), while play a supporting role in anti-inflammatory process (Pan et al., 2017) when expressed on cell surface of vascular endothelial cells. P2RY8, an orphan G protein-coupled receptor like GPR15, is a key regulator for affinity maturation of B cells in germinal centers (Muppidi et al., 2014). SNAI3 was identified as an invasion-related marker (Puisieux et al., 2014), and also reported to play a key role in differentiation of lymphoid cells and myeloid cells (Dahlem et al., 2012). TLR7, a dual receptor for guanosine and uridine-containing ssRNA in innate immunity, could perform tumor-suppressive activity by mediating the activation of NfκB and inducing proinflammatory cytokines (Schon and Schon, 2008). This gene has been recognized as a hot target for tumor targeted immunotherapy and cancer vaccines (Schon and Schon, 2008; Lynn et al., 2020). ATOH1 is a master TF for regeneration and differentiation of intestinal epithelial cells (Ishibashi et al., 2018). SIGLEC1, also named as CD169, could function as a facilitator of the recognition and internalization of sialic acid decorated apoptotic bodies and exosomes derived from tumors, which potentially contributes to both attenuation and facilitation of anti-tumor immunity (Fraschilla and Pillai, 2017). It is appealing that several genes among 4-DysReg possess immune-related functions, including RUNX3, GPR15, P2RY8, SNAI3, TLR7, and SIGLEC1. We have actually observed the positive correlation between the expression of these genes and the abundance of eight immune cell types within tumor tissue (Supplementary File and Supplementary Figure S5), which provides functional interpretability of 4-DysReg from the viewpoint of cellular functions among tissue. It is noted that some signature genes, e.g. RUNX3, display opposite functions in different studies. Although RUNX3 was reported to be a tumor suppressor in CRC (Weisenberger et al., 2006; Soong et al., 2009), in TCGA CRC dataset, RUNX3 expression displays weak difference between normal and cancer (Supplementary Figure S6) and weak correlation with prognosis (HR = 1.125, 95% CI: 0.965–1.312). In line with previous observations, it is suggested that cytoplasmic localization of RUNX3 is a main mode of RUNX3 inactivation (Soong et al., 2009; Ito et al., 2015), which explains why only RUNX3 displays limited prediction accuracy for prognosis (Figure 5A and C). Benefitting from gene dysregulation analysis, we obtained functional relevance of RUNX3 from its one target, GPR15. Still, in TCGA CRC dataset, the expression of GPR15 significantly decreases in cancer (Supplementary Figure S6) and is significantly correlated with good prognosis (HR = 0.735, 95% CI:...
of RUNX et al., post-transcriptional modification and TF combination (Lambert et al., 2016) and some accessibility, histone modification, DNA methylation, TF binding is also influenced by other features, such as chromatin. This implication is not limited to RUNX but also could be extended to other genes or regulations in the field. In the present study, RUNX provides a promising solution to address this issue. In the context of differential regulation analysis, or dysregulation analysis, we believe that the gene dysregulation analysis framework will help to elucidate systematic mechanisms of carcinogenesis and develop signatures with high predictive accuracy and mechanistic interpretability in clinical application.

**Materials and methods**

*Collecting and preprocessing expression data*

TCGA CRC mRNA data and clinical data were downloaded from UCSC Xena (Vivian et al., 2017), and 380 primary tumor samples and 51 adjacent normal samples were selected. The abundance of mRNA was determined as transcripts per million (TPM). The TPM values <1 were treated as missing values. For a certain gene, when the number of missing values was >20% of total sample size, the gene was deleted. The remaining missing data were filled in with k-nearest neighbor method (Troyanskaya et al., 2001). The expression data were log2 transformed. Among them, 32 pairs of matched tumor and adjacent normal samples were used to construct conditional GRNs and identify gene dysregulations; 350 primary tumor samples with OS, RFS, age, gender, and pathological stage information were used in prognostic analysis.

Two expression datasets with clinical data including OS, RFS, age, gender, and pathological stage, GSE39582 (Marisa et al., 2013) and GSE17538 (Smith et al., 2010), were downloaded from GEO (http://www.ncbi.nlm.nih.gov/geo/). GSE39582 contains 566 CRC tumor samples and 19 normal samples. GSE17538 contains 238 CRC tumor samples. Both datasets are based on GPL570 platform. The probes matching to multiple genes were deleted. When >1 probe could be mapped to the same gene, the probe with the highest value was taken to represent the expression level of the gene. The way of handling missing data was the same as that for TCGA data. The expression data were log2 transformed. Quantile method was adopted to normalize data between samples (Bolstad et al., 2003). A total of 553 tumor samples in GSE39582 and 200
tumor samples in GSE17538 with sufficient clinical information were used in prognostic analysis.

**Constructing reference GRN**

TF and the 'best' motif data were accessed from HumanTF database (Lambert et al., 2018). The 1000 bp upstream sequences from transcription start site of all coding genes were acquired by biomart (http://www.biomart.org/) and regarded as promoter regions. The promoter region of each gene was scanned with FIMO (Grant et al., 2011) with P-value < 1e–04. For each TF, the threshold that the probability of making at least one false discovery was < 0.01 was adopted to perform multi-hypothesis test correction to select its targets. If a TF had >5000 targets, only the top 5000 targets ranked by FIMO score were kept (Grant et al., 2011). The remaining TF–target relationships, or candidate TF–targets, composed the reference GRN (Figure 1A), which included 1083 TFs and 1773407 relationships.

**Constructing conditional GRNs**

Based on expression data under a specific condition, TF–target relationships in reference GRN were filtered by using a feature selection method Boruta, a wrapper around random forest algorithm (Kursa and Rudnicki, 2010). For each target, the expression values of its candidate TFs in reference GRN were regarded as original features, and the target's expression was regarded as response variable. By shuffling candidate TFs, Boruta generated shadow features, iteratively estimated the importance of every TF, and removed the TF less relevant to its target. In this way, those conceptual links which did not perform functions under specific condition were removed from reference GRN, and the remaining TF–target relationships formed conditional GRNs (Figure 1B), i.e. normal GRN and cancer GRN in the present study.

**Identifying gene dysregulations**

Gene dysregulations were identified by integrating three standards. Firstly, the regulatory intensity should be significantly different between conditions; secondly, target should be differentially expressed; lastly, the change direction of regulatory intensity should be consistent with the change direction of target's expression between conditions (Figure 1C).

Since the de-biased LASSO method is able to robustly obtain the estimation and the covariance of regression coefficients for high-dimensional regression (Javanmard and Montanari, 2014), it was adopted to estimate the regression coefficients and their CIs of every upstream TF of a certain target, with regression coefficients and CIs taken as regulatory intensities and their ranges. For each regulation, if 95% CIs of the regression coefficient have no overlap between normal and cancer conditions, the regulatory intensity could be regarded as significantly differential.

For target expression change, we carried out differential expression analysis by using limma (Ritchie et al., 2015) with the cutoff of |logFC| > 1 and $P_{adj} < 0.05$ on log2-transformed expression data. DEGs were identified and their change directions, say activation of inhibition, were recorded.

Eventually, for a certain regulation, the TF whose regulatory intensity change was consistent with its target's expression change from one condition to another was considered to play key role in controlling target's expression. And the regulations with consistent change of regulatory intensity and target's expression were kept in the following analysis.

**Checking prognostic effect of dysregulations**

For each gene dysregulation, four types of cox proportional hazard models were constructed in TCGA CRC dataset by using R package survival (https://cran.r-project.org/web/packages/survival/), fitted with (i) two genes' expression data adjuvant with clinical information including age, gender, and pathological stage for OS; (ii) two genes' expression data for OS, (iii) two genes' expression data adjuvant with clinical information including age, gender, and pathological stage for RFS, and (iv) two genes' expression data for RFS. For each cox model, C-index was calculated to measure the prediction accuracy. Meanwhile, four groups of controls were set with the following procedures, each containing the same number of randomly selected gene pairs as the identified gene dysregulations: (i) two genes were randomly selected from the gene list of preprocessed expression data; (ii) one DEG and one non-differentially expressed gene (non-DEG) were randomly selected from the gene list of preprocessed expression data; (iii) one regulation was randomly selected from the reference GRN; and (iv) one regulation whose target is DEG was randomly selected from the reference GRN. Still, four types of cox models were built for each random pair, and C-index was calculated for each model. At last, for each type of cox models, Wilcox test was used to check whether the C-indexes from the identified dysregulations were significantly larger than those from random pairs (one-way Wilcox test), i.e. whether the accuracy of cox models fitted by the identified dysregulations as a whole was higher than by controls. This process was repeated 100 times. The prognostic effect of the identified dysregulations was also validated in GSE39582 and GSE17538.

**Building prognostic signature**

Among the OS cox models fitted with gene expression data of every dysregulation adjuvant with clinical information in TCGA CRC dataset as described above, the model with the largest C-index was taken as the primary one, and then a greedy strategy was adopted to exclusively add every other candidate dysregulation with clinical information including age, gender, and pathological stage for OS, (ii) two genes' expression data for OS, (iii) two genes' expression data adjuvant with clinical information including age, gender, and pathological stage for RFS, and (iv) two genes' expression data for RFS. For each cox model, C-index was calculated to measure the prediction accuracy. Eventually, for a certain regulation, the TF whose regulatory intensity change was consistent with its target's expression change from one condition to another was considered to play key role in controlling target's expression. And the regulations with consistent change of regulatory intensity and target's expression were kept in the following analysis.

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Risk score of each sample was calculated by genes in the prognostic signature and clinical factors including age, gender, and histological type with cox multivariate regression among all samples in the dataset. Samples were divided into two groups with the median of risk scores, and Kaplan–Meier (KM) survival analysis was then carried out to evaluate the difference in survival time between the two groups. The accuracy of the risk score for predicting survival was evaluated by C-index and AUC with timeROC (Blanche et al., 2013). These analyses were also conducted in two independent datasets.

The prognostic signature was further validated following a cross-validation procedure. That is, 60% of the samples were randomly selected as training set to fit a cox model as described above, and the left 40% were taken as testing set to calculate C-index, AUC at 1-, 3-, 5-year OS, and significance of KM survival curves. The cross-validation was repeated 100 times. This cross-validation procedure was respectively conducted in TCGA CRC and two independent datasets.

**Enrichment analysis**

Pathway over-representation analysis was implemented on ConcensusPathDB (Herwig et al., 2016), which is a comprehensive collection of human molecular interaction data integrating different public repositories. The pathway databases used in ConcensusPathDB include KEGG, Reactome, WikiPathways, Biocarta, and PharmGKB. Over-representation pathways were extracted with the threshold that pathway had at least five genes overlapped with input genes and P-value was <0.05.

GSEA was performed with GSEA3.0 software (Subramanian et al., 2005). The gene set database ‘c5.bp.v6.2.symbols.gmt [Gene ontology]’ was used to determine the enrichment of gene sets.

**Supplementary material**

Supplementary material is available at *Journal of Molecular Cell Biology* online.

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