Effect of proton pump inhibitors on bone mineral density: A systematic review and meta-analysis of observational studies

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**Abstract**

Despite the fact that proton pump inhibitors (PPIs) are one of the most prescribed medications, several epidemiological studies have reported many adverse effects related to their long-term usage. Nevertheless, there were inconsistent findings in the literature with regard to PPI use and bone mineral density (BMD) change. The aim of this systematic review and meta-analysis is to evaluate the association between the use of PPIs and change in BMD. The PubMed/MEDLINE, EMBASE, Cochrane and CINAHL databases were searched up to March 2019. Ten studies fulfilled the eligibility criteria (4761 cases and 30,809 controls), from which the mean difference and mean annualized percent change in BMD were pooled using RevMan 5.3.5. The results showed no statistically significant association between PPI users and non-users in mean annualized percent change in BMD (0.06; 95% CI –0.07, 0.18) with moderate heterogeneity (I²: 63%). There was a statistically significant reduction in the mean BMD difference among PPI users (–0.03; 95% CI –0.04, –0.01) with no substantial heterogeneity (I²: 26%). This meta-analysis reported inconsistent results regarding the use of PPIs and BMD loss. Thus, the effect of PPIs on BMD needs to be elucidated by other studies, and healthcare providers should prescribe PPIs with caution considering their unfavorable consequences on bone health.

1. Introduction

Proton pump inhibitors (PPIs) were first introduced in 1989; since then, they have been commonly used for treating acid-related disorders such as gastroesophageal reflux disease (GERD) and peptic ulcer disease (PUD) (Strand et al., 2017; Hershcovici and Fass, 2010; Sandhu and Fass, 2018). PPIs are used worldwide as prescribed medications, and in many countries, they are available over-the-counter (OTC) medications due to their efficacy and known safety profile (Strand et al., 2017; Hershcovici and Fass, 2010). PPIs are used worldwide as prescribed medications, and in many countries, they are available over-the-counter (OTC) medications due to their efficacy and known safety profile (Strand et al., 2017; Hershcovici and Fass, 2010).

Recently, several studies reported many adverse effects related to the long-term use of PPIs, such as an increased risk of osteoporotic-related fractures, *Clostridium difficile* infection, pneumonia, and vitamin B12 and magnesium deficiencies (Heidelbaugh, 2013; Nehra et al., 2018; Wang et al., 2019; Savarino et al., 2018; Pezeshkian and Conway, 2018).

Many systematic reviews and observational studies have shown an increased risk of osteoporotic fractures after the long-term use of PPIs (Zhou et al., 2016; Nassar and Richter, 2018; Hussain et al., 2018; Islam et al., 2018; Abramowitz et al., 2016). However, the precise mechanism remains unclear, and the causality of the association is inconclusive. Studies that assessed the relationship between PPI use and increased osteoporotic fracture risk suggested several mechanisms, such as a reduction in intestinal calcium absorption, an interruption in osteoclast function in bone remolding and repair, and a decrease in bone mineral density (BMD) (Ito and Jensen, 2010; Arj et al., 2016; Maleth and Hegyi, 2013).

Although many epidemiological studies reported that PPI treatment reduces BMD (Heidelbaugh, 2013; Lau and Ahmed, 2012), others failed to find a significant association (Lau and Ahmed, 2012). Moreover, two systematic reviews and meta-analyses with different inclusion criteria reported no significant difference in the mean values of BMD between PPI users and controls (Zhou et al., 2016; Nassar and Richter, 2018). These results attract our attention to conduct a systematic review and meta-analysis of existing observational studies to evaluate the association between the use of PPIs and changes in BMD.
2. Methodology

This systematic review and meta-analysis was conducted following the preferred reporting items for systematic reviews and meta-analyses (PRISMA) guidelines (Moher et al., 2009) to explore the association of PPI use and BMD change.

2.1. Eligibility criteria

All studies that fulfilled the following criteria were included: (a) cohort or case-control study design; (b) study population above 18 years of age in both sexes; (c) the use of PPI was defined as an exposure; (d) the change in BMD was reported as an outcome by using dual-energy X-ray absorptiometry (DXA or DEXA); and (e) the means ± standard deviations (SDs) were provided for the PPI users and the control group or adequate information was provided to calculate them. Studies were excluded if they have the following criteria: (a) the study examined the association between PPI use and change in BMD in combination with histamine2-receptor antagonists or other drugs that affect bone metabolism such as bisphosphonate or glucocorticoids and/or (b) use of peripheral quantitative computed tomography scans (pQCT) or other methods rather than DXA or DEXA for the measurement of BMD.

2.2. Search strategy

A comprehensive literature search was performed up to March 2019 in electronic databases including the PubMed/MEDLINE (national center for biotechnology information), EMBASE (Elsevier), Cochrane (Wiley online library) and Cumulative Index to Nursing and Allied Health Literature (CINAHL) databases without restriction to language using the following keywords: (a) proton pump inhibitors, lansoprazole, pantoprazole, rabeprazole, esomeprazole, omeprazole, dexlansoprazole, gastric acid-suppressive agents, gastric acid inhibitors, antacid, OR antiulcer agents; (b) osteoporosis, bone mineral density OR osteopenia. Moreover, a manual search of the retrieved articles’ references was conducted.

2.3. Study selection and data extraction

Studies that satisfied the inclusion criteria were selected by two independent reviewers (SE and SH), and conflicts were resolved by a third investigator (MF). The two reviewers abstracted the qualitative and quantitative data from the included articles by using a designed data extraction template including the following: study author, year of publication, study country/settlement, study design, study period, study population sex/mean age/number of controls/numbers of PPI users, exposure type/dose/duration, outcome mean ± SD and P value.

2.4. Risk of bias and quality assessment

An assessment of the quality of the included cohort studies was performed by two independent reviewers (SE and SH) using the Newcastle–Ottawa scale (NOS) (Wells et al., 2015) for evaluating the quality of nonrandomized studies in meta-analyses. Three factors were considered for scoring the quality of the studies: (1) selection, (2) comparability, and (3) outcome (Wells et al., 2015). The quality of the studies was rated by awarding stars in each subset with a total maximum score of 9 (Wells et al., 2015). Studies that scored ≥7 were considered high-quality, while those that scored <7 were considered low-quality.

2.5. Statistical analysis

Statistical analysis was performed using Review Manager version 5.3.5 (RevMan 5.3.5) (Review Manager (RevMan), 2014) to assess the association between PPI use and change in mean or annualized mean percent change in BMD. A fixed effects model was used for the meta-analysis in the absence of substantial heterogeneity, and a random effects model was used for the meta-analysis when heterogeneity existed. Furthermore, the I2 level was calculated to evaluate the heterogeneity across studies. The I2 value lies between 0 and 100%, and an I2 value of 75% represents a high level of heterogeneity; likewise, 50% represents a substantial level of heterogeneity. Subgroup analysis was performed based on body site (hip, femoral neck, and spine). However, subgrouping was not performed based on PPI types due to insufficient data for comparison in the selected studies.

2.6. Dealing with missing data

The SDs of the mean difference before and after treatment were missing in some of the selected studies. Therefore, the authors of these studies were contacted, but the data were not retrieved; thus, the missing data were recalculated by a statistician (IA). The formula for calculation of standard deviation is: SD = \sqrt{\frac{\sum (f_x)^2 - (\sum f_x)^2}{N}}

2.7. Ethical approval

The present study was approved by the King Abdullah International Medical Research Center IRB committee on Feb 25, 2018 (research number RC18/048/R) (Mustafa, n.d.).

3. Results

3.1. Literature search

The search strategy identified a total of 853 citations. After removing duplicates, 772 relevant studies remained. Based on the title and abstract screening, 717 articles were excluded. Out of 55 full-text articles, 10 observational studies met the inclusion criteria for qualitative and quantitative analyses (Gray et al., 2010; Solomon et al., 2015; Bahtiri et al., 2016; Targownik et al., 2017; Shin et al., 2019; Xuan et al., 2014; Ozdil et al., 2013; Roux et al., 2012; Targownik et al., 2012; Elaine et al., 2008; Elaine et al., 2011) (Fig. 1).

3.2. Study characteristics

In the literature search, 10 longitudinal cohort design studies reported the association between PPI use and change in BMD. The retrieved studies were published between 2008 and 2018. Nine studies were in English, and one was in Chinese. The studies were conducted in the USA, the Republic of Kosovo, Canada, China, Turkey and South Korea (Gray et al., 2010; Solomon et al., 2015; Bahtiri et al., 2016; Targownik et al., 2017; Shin et al., 2019; Xuan et al., 2014; Ozdil et al., 2013; Roux et al., 2012; Targownik et al., 2012; Elaine et al., 2008). Overall, the data for 4761 cases and 30,809 controls including males and females with ages ranging from 28.9 to 83.1 years were used in this meta-analysis. Different PPI classes were used with a range of duration of 30 days to 10 years. In the included studies, BMD change was measured by DXA using Hologic instruments. Most of the studies’ results were adjusted for age, sex, and BMI (Table 1).

3.3. Primary outcomes

Two studies provided mean annualized percent change in BMD between PPI users and nonusers, and four studies provided mean differences in BMD. Studies with mean difference in BMD results favored no PPI usage with mean difference of −0.03 (95% CI −0.04, 0.01), and no substantial heterogeneity (I2: 26%) was found. Among the studies reported, the mean annualized percent change in BMD results showed no statistically significant association between PPI users and nonusers with a mean difference of 0.06 (95% CI −0.07, 0.18) and moderate heterogeneity (I2: 63%) (Fig. 2).
3.4. Subgroup analysis

Subgroup analysis was performed based on anatomic sites. The mean difference in BMD was statistically significant in both the spine and femur BMD results: hip $0.02$ (95% CI $0.04, 0.00$) with moderate heterogeneity ($I^2$: 54%); spine $0.03$ (95% CI $0.06, 0.01$) with no substantial heterogeneity ($I^2$: 26%); and femur $0.04$ (95% CI $0.06, 0.01$) with no heterogeneity ($I^2$: 0%). For the mean annualized percent change in BMD, there was no statistical significance as follows: hip $0.24$ (95% CI $0.42, 0.90$) with substantial heterogeneity ($I^2$: 89%); spine $0.04$ (95% CI $0.45, 0.38$) with moderate heterogeneity ($I^2$: 52%); and femur $0.07$ (95% CI $0.01, 0.13$) with no heterogeneity ($I^2$: 0%) (Fig. 3).

4. Discussion

This systematic review and meta-analysis of cohort studies that aimed to explore the association between PPI usage and the change in BMD revealed a reduction in BMD among PPI users regarding the mean difference in BMD. In contrast, there was no statistically significant association between PPI users and nonusers regarding the mean annualized percent change in BMD.

Many systematic reviews have reported an association between PPI use and an increased risk of fractures (Zhou et al., 2016; Elaine et al., 2011). However, the exact underlying mechanism remains unclear (Zhou et al., 2016; Elaine et al., 2011). One of the potential mechanisms that has been suggested in theory is that the role of PPIs in blocking gastric acid secretion could reduce intestinal calcium absorption and lead to a decrease in BMD (Ito and Jensen, 2010). Two recently published systematic reviews that addressed this question reported no significant difference in BMD between PPI users and nonusers (Zhou et al., 2016; Nassar and Richter, 2018), supporting the results of the mean annualized percent change in BMD in this study. Liu et al. failed to find a correlation between PPI use and BMD loss in the femur (SMD: $0.27; 95\% \text{ CI} -0.62, 0.09$) and spine (SMD: $-0.06; 95\% \text{ CI} -0.54, 0.41$), but the results were from only three studies (Zhou et al., 2016). Additionally, Nassar et al. showed no significant difference in the standardized mean differences in BMD between PPI users and controls in cross-sectional (SMD: $0.00; 95\% \text{ CI} -0.18, 0.19$) and longitudinal BMD values (SMD: $0.07; 95\% \text{ CI} -0.06, 0.20$) (Nassar and Richter, 2018). Several articles included in the abovementioned systematic reviews and meta-analyses are not included in this study for two reasons. The first is the difficulty of inferring the temporal association between PPI use and change in BMD in cross-sectional studies. The second is the inability to mask the effect of bisphosphonates on bone health among the population using bisphosphonates simultaneously with PPIs.

Several observational studies showed conflicting results with regard to the PPI risk of low BMD. A prospective study by Yu et al. in men and women over the age of 65 demonstrated no statistically significant differences in the effect of the long-term use of PPI on BMD after adjusting...
for potential confounding factors, including age, race, BMI, alcohol use, exercise, oral or inhaled corticosteroid use, NSAID use, calcium supplement use, osteoporosis medication use, and self-reported health and concurrent weight change (Elaine et al., 2008). Furthermore, Targownik et al. observed no association between continuous PPI use and the rate of change in BMD at different measurement sites (including the total hip, femoral neck, and lumbar spine) between PPI users and nonusers at the 5-year and 10-year follow-ups (Targownik et al., 2012). Nevertheless, Ozdil et al. revealed statistically significant reductions in densitometric T-scores of the vertebra and femur in patients who used PPIs (Ozdil et al., 2013). Considering that the mean age group was 37.7 ± 8.8 and the duration of exposure was 8.5 ± 2.3 months, this population was the youngest and had the shortest duration compared to the studies mentioned above.

Most of the existing epidemiological evidence investigated the relationship between PPI use and BMD using DXA, which is a two-dimensional technique that provides measurements of areal (mg/cm2) rather than volumetric (mg/cm3) BMD (Gray et al., 2010; Bahtiri et al., 2015). As a result, these studies by using DXA failed to assess the independent changes of PPI on trabecular and cortical compartment (Targownik et al., 2017; Stathopoulos et al., 2016; Lauretani et al., 2006; Russo et al., 2003). A study by Maggio et al. investigated the relationship between the chronic use of PPIs and the cortical and trabecular BMD (vBMDc and vBMDt, respectively) in older individuals by using Tibial pQCT scans; their data showed a negative association between PPI use and vBMD in a small sample of community-dwelling older persons (Marcello Maggio et al., 2013). These findings support the hypothesis of a possible direct effect of PPIs on bone mineral metabolism despite the no significant association between the use of PPIs and bone geometry.

Regardless of the contradictory results concerning the effect of PPIs on BMD, many systematic reviews revealed an association between PPI use and increased fracture incidence (Zhou et al., 2016; Elaine et al., 2011). Thus, taking PPIs for inappropriate indications or receiving PPI therapy without prescription should be discouraged. Additionally, the negative influence of PPIs on bone quality and bone mineral metabolism should be explored by further studies with larger sample sizes and longer durations, considering other possible confounding factors and different settings to resolve the dispute.

This meta-analysis has a number of remarkable strengths, such as the strict inclusion criteria that focused on only the relationship between PPI use and BMD change. Another strength is the inclusion of observational studies that were published with no language restrictions. Nevertheless, there are several limitations, some of which are the involvement of few...
observational studies due to the paucity of randomized controlled trials. Therefore, the possibility of residual confounding from various unmeasured variables cannot be excluded. Additionally, significant heterogeneity was present in the outcomes of the studies that provided the mean annualized percent change in BMD. This heterogeneity might be explained by a variety of exposure durations, types of PPI use, and the sex, mean age and sample size of the study populations. Accordingly, the results should be interpreted with caution to better explain the significant statistical and clinical heterogeneity among the studies, and the inability of observational studies to clarify whether the observed association is a causal effect or a result of unmeasured variables should be considered.

5. Conclusion

The literature reviewed that evaluated the risk of PPI use and BMD loss convey inconsistent results with BMD reduction among studies with mean differences and no BMD change among studies with annualized

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Fig. 2. Forest plot showing the meta-analysis of mean annualized percent change in bone mineral density between proton-pump inhibitor (PPI) users and nonusers in the hip, spine & femur.

Fig. 3. Forest plot showing the meta-analysis of mean differences in bone mineral density between proton-pump inhibitor (PPI) users and nonusers in the hip, spine & femur.
means. Therefore, well-conducted observational and randomized controlled trials are recommended to resolve the conflict. Moreover, healthcare providers should consider the potentially unfavorable effect of PPI use on bone health.

CRediT authorship contribution statement

Research idea by SA and SH. Comprehensive search, abstraction, selection of articles, and manuscript writing by SA and SH. Supervision and resolution of any conflicts in each step by MF.

Critical review and editing by SA and SH.

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Declaration of competing interest

We declare that there are no conflicts of interest, and we confirm that this manuscript has not been published elsewhere and is not under consideration by another journal.

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