Variational ansatz for gaussian + Yang-Mills two matrix model compared with Monte-Carlo simulations in ’t Hooft limit

Govind S. Krishnaswami
Department of Physics and Astronomy,
University of Rochester, Rochester, New York 14627

March 27, 2022

Abstract

In recent work, we have developed a variational principle for large \(N\) multi-matrix models based on the extremization of non-commutative entropy. Here, we test the simplest variational ansatz for our entropic variational principle with Monte-Carlo measurements. In particular, we study the two matrix model with action \(\text{tr} \left[ \frac{m^2}{2}(A_1^2 + A_2^2) - \frac{i}{2}[A_1, A_2]^2 \right]\) which has not been exactly solved. We estimate the expectation values of traces of products of matrices and also those of traces of products of exponentials of matrices (Wilson loop operators). These are compared with a Monte-Carlo simulation. We find that the simplest wignerian variational ansatz provides a remarkably good estimate for observables when \(m^2\) is of order unity or more. For small values of \(m^2\) the wignerian ansatz is not a good approximation: the measured correlations grow without bound, reflecting the non-convergence of matrix integrals defining the pure commutator squared action. Comparison of this ansatz with the exact solution of a two matrix model studied by Mehta is also summarized. Here the wignerian ansatz is a good approximation both for strong and weak coupling.

KEYWORDS: multi matrix models, Yang-Mills integrals, Yang-Mills theory, M(atrix) theory, large \(N\) limit, entropy, variational principle, Monte-Carlo integration.
1 Introduction

Matrix models are of long standing interest in several branches of physics and mathematics.

Early work of Wigner, Dyson and others introduced random matrices in the study of statistical properties of highly excited energy levels of nuclei [1].

In the early 1970s the work of ’t Hooft [2, 3] on the large $N$ limit of QCD ($N$ is the number of colors, gluon fields are $N \times N$ matrices), made the study of large $N$ matrix field theories a central theme in understanding the non-perturbative dynamics of non-abelian gauge theories. This also gave the first indication that it is in the large $N$ limit that a gauge theory may have the long sought dual description as a string theory.

Important progress was made in the late 1970s and early 1980s stemming from Migdal and Makeenko’s work on the factorized loop equations of large $N$ gauge theories [4]. The work of Sakita and Jevicki [5] on the collective field formalism of large $N$ field theories and that of Cvitanovic and collaborators [6] also bears mention from this period. Eguchi and Kawai’s [7] proposal on reducing a matrix field theory to a matrix model with a finite number of degrees of freedom but in the large $N$ limit has been a recurring theme ever since.

Important breakthroughs in the study of random surfaces, two dimensional string theory and two dimensional gravity coupled to matter were made in the late 80s and early 1990s (see [8] for a review) The planar Feynman graph expansion of large $N$ matrix models was used as a way of descretizing a two dimensional surface. Models with one or a finite number of matrices and the $c = 1$ quantum mechanics of a single matrix were of importance in these developments. In addition to the large $N$ limit, the double scaling limit was developed to study the surfaces obtained in the continuum limit.

From the early 1990s onwards, the work of mathematicians including Voiculescu and collaborators on von Neumann algebras lead to the development of the field of non-commutative probability theory [9]. Large $N$ matrix models are natural examples of non-commutative probability theories. Random matrices also have deep connections to the statistical properties of zeros of the Riemann zeta function [10].

In the mid 1990s, supersymmetric matrix models we proposed as non-perturbative
definitions of M-theory and superstring theory [10, 11]. Bosonic matrix models are also studied [12, 13, 14, 15, 16, 17] as a first step towards understanding these supersymmetric matrix models.

Recently, interest in matrix models has been revived in several contexts.

Work of Dijkgraaf and Vafa [18, 19] has shown that the effective superpotential of $\mathcal{N} = 1$ supersymmetric gauge theories with an adjoint chiral superfield can be calculated exactly from a bosonic one matrix model in the large $N$ limit. This correspondence also has extensions to multi-matrix models, for example the glueball superpotential of $\mathcal{N} = 1^*$ gauge theory was determined from the partition function of the corresponding three-matrix model [19, 22, 21]. Bosonic multi-matrix models also arise in the context of quiver gauge theories [20].

The $c = 1$ matrix model has been revived in the context of two dimensional super-gravity coupled to $\hat{c} = 1$ matter [23]. $c = 1$ matrix models are also of current interest in studying unstable D-branes and the phenomenon of tachyon condensation [24, 25].

We may also regard finite matrix models as zero momentum limits of matrix field theories such as Yang-Mills theory. They provide a testing ground for new ideas on matrix field theories.

While matrix models have been extending their influence into a wide variety of contexts, there have been attempts to solve matrix models exactly. For instance, the work of Brezin et.al. [26], Mehta [27], Kazakov [29], Staudacher [28] and others has shed light on the partition functions and certain special classes of correlations in one and two matrix models. The two matrix model studied by Mehta is actually a special case of a class of chain multi-matrix models which can all be solved exactly, leading in the continuum limit to the solution of the $c = 1$ matrix quantum mechanics found by Brezin et. al.

However, it has proven difficult to determine the correlations of a generic multi-matrix model analytically. In the ’t Hooft limit the fluctuations in invariant observables becomes small and the theory becomes classical while retaining the quantum fluctuations in $\hbar$ and certain other non-perturbative features. Aside from some special cases where exact solutions for the partition function and certain correlations have been possible, multi-matrix models, even in the large $N$ limit are not exactly solved. This should not be surprising, since they
represent complicated classical dynamical systems. Monte-carlo integration is probably the best available means of obtaining numerically accurate predictions. Their drawback is that some of the mathematical structures are not revealed. It would be useful to have a middle ground: an approximation method that provides both qualitative and mathematical insights and a quantitative estimate for correlations.

In a previous paper ([30], see also [31, 32]) we derived a variational principle for large $N$ multi-matrix models. This involved many new theoretical ideas, especially the role played by the automorphism group of the free algebra generated by the matrices. The most important physical principle learnt was that the cohomology of this group is a non-commutative entropy. It was the crucial element in deriving a classical action for large $N$ matrix models. This allowed us to obtain variational approximations for the correlation tensors of multi-matrix models in the large $N$ limit. Thus we have a self contained formulation and method of approximate solution for matrix models in the $N \rightarrow \infty$ limit. It deals directly with correlations, rather than matrix elements. There is no more any reference to integrations over matrices nor the principle of unitary invariance. Rather, we have a classical theory on the configuration space of non-commutative probability distributions. The correlation tensors are coordinates on this space. They are determined by a constrained maximization of a non-commutative analogue of entropy. The action of the original matrix model is encoded in the constraints.

To solve this extremization problem approximately, we maximize the entropy on a conveniently chosen finite-dimensional sub-space of the configuration space. The simplest such choice is the subspace corresponding to the wignerian correlations.

In [30] we compared our variational ansatz with the exact solution of a two matrix model studied by Mehta [27]. The variational ansatz gave a reasonably good approximation both for strong and weak coupling (summarized in section 5). This gives us the confidence to test our ansatz for models that have not been exactly solved. In this paper we quantitatively test the wignerian variational ansatz with correlations measured using Monte-Carlo integration for a bosonic two matrix model with action $\text{tr} \left[ \frac{1}{2} m^2 (A_1^2 + A_2^2) - \frac{1}{4} [A_1, A_2]^2 \right]$. We
study a two matrix model since it is the simplest multi-matrix model. We pick this action since it mimics that of the zero momentum limit of Yang-Mills theory. Moreover, this is the simplest two matrix model that is not exactly solved and also shares the derivation property of Yang-Mills theory. We use the Metropolis algorithm implemented on a personal computer to test our new variational principle with the simplest of ansatze. We compare measured and variational two and four point correlations (eg. \( < \frac{1}{4} [A_1 A_2 A_2 A_1] > \)) and also the expectation values of Wilson loop operators (eg. \( < \frac{i}{2} e^{ilA_1} e^{ilA_2} e^{-ilA_1} e^{-ilA_2} > \)) in the large N limit. We find that it works remarkably well for \( m^2 \) of order 1 or more. As \( m^2 \to 0 \), the measured correlations appear to grow without bound, reflecting the divergence in the matrix integrals that define the pure commutator squared model. This is not captured by the wignerian ansatz.

We now summarize the framework we use to study large N matrix models.

2 Large N Matrix Models

We consider multi-matrix models where the dynamical variables are a set of \( M \) hermitian \( N \times N \) matrices \([A_i]_{ab}^a\). Here \( i = 1, \cdots, M \) labels the matrices and \( a, b = 1, \cdots, N \) are the ‘color’ row and column indices. The action \( S(A) \) for the matrix model is a \( U(N) \) invariant polynomial in \( S(A) = tr S^I A_I \). \(^1\) An example is the 2 matrix model

\[
S = tr \left[ -\frac{1}{4} [A_1, A_2]^2 \right]
\]

Let

\[
Z = \int dA e^{-NS(A)}
\]

denote the partition function. The observables we are interested in are the correlation tensors \( G_I \) of the large N limit, the limit where \( N \to \infty \) holding the coupling constants \( S^I \) fixed.

\[
< \Phi_{i_1 \cdots i_p} > \equiv < \frac{1}{N} A_{i_1} \cdots A_{i_p} > = \frac{1}{Z} \int dA e^{-NS(A)} \frac{1}{N} A_{i_1} \cdots A_{i_p}
\]

\(^1\)Capital letters denote multi-indices \( S^I = S^{i_1 \cdots i_p} \), \( A_I = A_{i_1} A_{i_2} \cdots A_{i_p} \)
\[ G_{i_1 \cdots i_p} = \lim_{N \to \infty} \langle \Phi_{i_1 \cdots i_p} \rangle \]  

These are a complete set of observables in the \( N \to \infty \) limit since expectation values of products of invariants factorize

\[ \langle \Phi_{I_1} \Phi_{I_2} \cdots \Phi_{I_r} \rangle = \langle \Phi_{I_1} \rangle \langle \Phi_{I_2} \rangle \cdots \langle \Phi_{I_r} \rangle + O\left( \frac{1}{N^2} \right) \]  

The \( G_I \) satisfy factorized loop equations (factorized Schwinger Dyson equations):

\[ S_{I_1 I_2 J_1 J_2} = \delta_{I_1 I_2} G_{J_1 I} G_{J_2} \]  

On the left side is an action dependent term while on the right is an anomalous universal term related to the non-commutative entropy explained in [30].

It has been shown [17] that the integrals over matrices for the pure commutator squared action for a two matrix model are not convergent. To see this, consider the partition function, and go to the basis in which \( A_1 \) is diagonal. In this basis, the integrand is independent of the diagonal elements of \( A_2 \), and therefore diverges. The divergence is even worse if we consider the expectation of the trace of a polynomial involving \( A_2 \), since then the integrand grows for large values of the diagonal elements of \( A_2 \). It is thus necessary for us to regularize the action. The simplest possibility is to add a quadratic term which ensures convergence of the matrix integrals:

\[ S = \frac{1}{2} (A_1^2 + A_2^2) - \frac{g^2}{4} [A_1, A_2]^2 \]  

This is the model we focus on. There is another reason to consider this two matrix model. An important property of Yang-Mills theory (in any dimension) is that its action leads to factorized loop equations whose action-dependent term is a derivation of the shuffle product of correlation tensors. Therefore we would like to study a two matrix model whose action shares this property. It can be shown that the most general quartic \(^3\) M-matrix model with this property is

\(^2\)This remark will be explained in detail in a forthcoming paper on the algebraic structure of factorized loop equations [33].

\(^3\)There are polynomial interactions of higher order with the derivation property.
\[ S = \text{tr} \left[ S_{i_1i_2}A_{i_1i_2} + S_{i_1i_2i_3i_4}[A_{i_1i_2}A_{i_3i_4} + S_{i_1i_2i_3i_4}(A_{i_1i_2i_3i_4} - A_{i_3i_1i_2i_4}) \right] \] (7)

where \( S_I \) are arbitrary cyclically symmetric tensors. We can use a \( GL_M(C) \) change of basis \( \tilde{A}_{i} = T_{i}^{j}A_{j} \) to reduce the action to the canonical form where the covariance \( S_{ij} \rightarrow \delta_{ij} \). Under such a change of basis \( S(A) = S_{i_1...i_n}A_{i_1...i_n} \rightarrow \tilde{S}(A) = S_{i_1...i_n}T_{i_1}^{j_1}...T_{i_n}^{j_n}A_{j_1...j_n} \) and \( G_{i_1...i_n} = T_{i_1}^{j_1}...T_{i_n}^{j_n}\tilde{G}_{j_1...j_n} \). Thus the correlations \( G_I \) can be obtained from those of the canonical action. In the two matrix case \( (M=2) \) the action reduces to

\[ S = \text{tr} \left[ \frac{1}{2}(A_1^2 + A_2^2) - \frac{g^2}{4}[A_1, A_2]^2 \right] \] (8)

There is only one independent coupling constant, the ratio of the coefficients of the quadratic and quartic terms. It is more convenient to study

\[ S = \text{tr} \left[ \frac{m^2}{2}(A_1^2 + A_2^2) - \frac{1}{4}[A_1, A_2]^2 \right] \] (9)

since it allows us to consider the pure commutator squared model in the \( m^2 \rightarrow 0 \) limit. In the large \( N \) limit, all correlations are functions of the single coupling constant \( m^2 \). In the pure commutator squared model, the coupling constant is an overall factor in the action and can be scaled out, the dependence of correlations on it can be determined by dimensional analysis. By contrast, in the model we study, the coupling constant dependence of the correlations is to be dynamically determined, making it more analogous to Yang-Mills theory than the large \( N \) reduced models of M-theory \cite{12, 13, 14, 15, 16, 17}.

Let us now summarize how we measure correlations by Monte-Carlo simulation.

### 3 Monte Carlo Measurement of Correlations

To measure the correlations numerically, we generate an ensemble of matrix configurations \( A_i^{(k)}, k = 1, \cdots, n \) such that as \( n \rightarrow \infty \), matrix elements \( [A_i]_{ab} \) picked at random from this ensemble are distributed according to \( \frac{1}{Z}e^{-NS(A)} \).
The Metropolis algorithm [34, 35, 36, 37] is used to create such an ensemble. We begin with a configuration \( A^{(1)} \) such as the zero matrices. The matrix elements are updated sequentially preserving hermiticity. At the \( k^{th} \) time step, a candidate configuration is generated \( B_i = A_i^{(k)} + W_i + W_i^\dagger \). Here for each \( i \), \( W_i \) is a random \( N \times N \) Weyl matrix. The only non-vanishing matrix element of \( W_i \) is picked at random from a uniform distribution of complex numbers in a square whose diagonally opposite vertices are at \(-\Lambda - \sqrt{-1}\Lambda, \Lambda + \sqrt{-1}\Lambda\).

If \( B \) has a greater Boltzmann weight than \( A^{(k)} \), the candidate is accepted,

\[
A_i^{(k+1)} = B_i \quad \text{if} \quad e^{-N(S(B) - S(A^{(k)}))} \geq 1
\]

and when \( B \) has a lesser Boltzmann weight than \( A^{(k)} \), the change is accepted with probability \( e^{-N(S(B) - S(A^{(k)}))} \) and rejected otherwise

\[
A_i^{(k+1)} = B_i \quad \text{if} \quad 1 \geq e^{-N(S(B) - S(A^{(k)}))} > r
\]

\[
A_i^{(k+1)} = A_i^{(k)} \quad \text{otherwise}
\]

Here \( r \) is a random number uniformly distributed in \((0, 1)\).

We can regard each matrix element as a continuous spin variable. One Metropolis sweep corresponds to \( \frac{1}{2}(N^2 - N) + N \) time steps of sequentially updating each independent matrix element of the \( M \) matrices. We perform a large number \( n_s \) of such Metropolis sweeps, generating an ensemble with \( n = \frac{1}{2}n_s(N^2 + N) \) configurations \((n = 18000\) for the measurements presented in this paper).

It is shown [35, 36] that as \( n \to \infty \) this algorithm produces a Boltzmann ensemble of configurations. The idea is that the above rules for making a transition can be used to define a Markov matrix on the space of ensembles. Its eigenvalue of maximum modulus \((=1)\) corresponds to the eigenvector labelling the Boltzmann ensemble. Thus one defines a contraction mapping on the space of ensembles whose unique fixed point is the Boltzmann ensemble.

Once we have this ensemble, the correlations are computed:

\[
< \frac{\text{tr}}{N} A_{i_1} \cdots A_{i_p} > = \frac{1}{n} \sum_{k=1}^{n} \frac{\text{tr}}{N} A_{i_1}^{(k)} \cdots A_{i_p}^{(k)}
\]
To extract the large $N$ limit of the correlations, we do the measurement for several values of $N = 10, 7, ..., 15$ and fit the results to the known $N$ dependence for large $N$:

$$< \frac{\text{tr}}{N} A_{i_1} \cdots A_{i_p} > \rightarrow G_{i_1 \cdots i_p} + \frac{\tilde{G}_{i_1 \cdots i_p}}{N^2} \quad (13)$$

and extract the value of $G_{i_1 \cdots i_p}$.

We mention the main sources of error in these measurements. First, there is the statistical error ($O(\frac{\text{variance}}{\sqrt{n}})$) of truncating the ensemble of configurations at a finite value of $n$. This is estimated by the bootstrap [38] procedure. Next there is the systematic error that could arise from the choice of initial configuration of matrices. This is estimated by changing the initial configuration slightly. The truncation ($\Lambda$) of the region in the complex plane from where the increments to random matrix elements are picked is a third source of error. As a practical matter we pick $\Lambda$ so that the acceptance of the algorithm to candidate configurations is roughly 50%. For sufficiently large $n$, we find the measurements to be insensitive to small changes in the value of $\Lambda$. The value of $\Lambda$ needs to be increased in order to measure correlations of very high order accurately while holding other parameters fixed. Finally, there is the error in extracting the large $N$ limit of the correlations from finite $N$ data.

4 Variational Principle for Large $N$ Matrix Models

Let us summarize the variational principle introduced in [30]. Given an action $S(A) = \text{tr} \, S^I A_I$, for an $M$ matrix model we want to determine the correlations $G_I$ in the large $N$ limit. We found a variational principle $\Omega(G) = \chi(G) - S^I G_I$ whose extremization leads to the factorized loop equations $S^I_{J_1 J_2} G_{J_1 J_2} = \delta^I_{J_1 J_2} G_{J_1 J_2}$. The variation of $S^I G_I$ gives the action dependent term on the left while the variation of $\chi$ gives the universal term on the right. The main difficulty was that the non-commutative entropy $\chi$ is not a power series in the $G_I$. The space of correlations is a coset space of the automorphism group of the free algebra by the subgroup of measure preserving automorphisms. So we expressed $\chi$ as an invariant power series on the larger space of automorphisms of
the free algebra. $\chi$ is actually a non-trivial 1-cocycle of this group. The formula for $\chi$ is given in [30], we will only use a special case of it in this paper.

Thus, to determine the correlations $G_J$, we must maximize the entropy $\chi$ while holding the correlations $G_I$ conjugate to the coupling tensors $S^I$ in the given action fixed. In other words, in the factorized loop equations, $S^I$ are the Lagrange multipliers enforcing these constraints.

The resulting maximum value $\chi_{\text{max}}$ has a simple physical meaning. Let $S_0(A) = \frac{1}{2} \text{tr} \sum_{i=1}^{M} A_i A_i$ be the canonical gaussian action for an $M$ matrix model. Suppose $Z$ and $Z_0$ are the partition functions of $S$ and $S_0$. Then

$$\chi_{\text{max}} = \lim_{N \to \infty} \frac{1}{N^2} \log \left[ \frac{Z}{Z_0} \right]$$

i.e. $-\chi_{\text{max}}$ is the free energy or vacuum energy measured with respect to the canonical gaussian matrix model.

Maximizing $\Omega$ exactly is equivalent to solving the factorized loop equations exactly, which has not been possible in general. To determine the correlations approximately, we maximize $\Omega$ on a subspace of the configuration space. The simplest subspace is that corresponding to the correlations of the multivariate Wigner distribution. This is the wignerian variational ansatz. On this subspace

$$\Omega[G_{ij}] = \frac{1}{2} \log \det[G_{ij}] - S^I G_I$$

where $G_I$ are the correlations of the multi-variate Wigner distribution. They are determined in terms of the two point correlation matrix by the planar analogue of Wick’s theorem: $G_{ijkl} = G_{ij} G_{kl} + G_{il} G_{jk}$, etc. Thus we regard the matrix elements of $G_{ij}$ as variational parameters and maximize $\Omega$ for the given action $S$.

To summarize, the wignerian variational ansatz gives the wignerian correlations that best approximate the true correlations of a given matrix model. Best approximation here means the one that maximizes entropy while holding the correlations conjugate to $S^I$ fixed.

We now compare our wignerian variational ansatz, first with the exact solution of a two matrix model studied by Mehta and then with the numerical solution of the gaussian + Yang-Mills two matrix model.
5 Two Matrix Model studied by Mehta

We recall here (for details see [30]) the comparison of our variational ansatz with Mehta’s exact solution of a two matrix model. In [27] Mehta finds the exact vacuum energy of the two matrix model with action

\[ S(A, B) = \text{tr} \left[ \frac{1}{2}(A^2 + B^2 - cAB - cBA) + \frac{g}{4}(A^4 + B^4) \right] \tag{16} \]

From his solution, we may extract the exact values of

\[ E^{ex}(g, c) = -\lim_{N \to \infty} \frac{1}{N^2} \log \frac{Z(g, c)}{Z(0, 0)}, \quad G^{ex}_{AB} \text{ and } G^{ex}_{AAAA}. \tag{17} \]

in the strong and weak coupling regimes. These are compared with our variational estimates below. For small \( g \) and \( c = \frac{1}{2} \):

\[
E^{ex}(g, \frac{1}{2}) = -1.44 + 1.78g - 8.74g^2 + \cdots \\
E^{var}(g, \frac{1}{2}) = -1.44 + 3.56g - 23.7g^2 + \cdots \\
G^{ex}_{AB}(g, \frac{1}{2}) = \frac{2}{3} - 4.74g + 53.33g^2 + \cdots \\
G^{var}_{AB}(g, \frac{1}{2}) = \frac{2}{3} - 4.74g + 48.46g^2 + \cdots \\
G^{ex}_{AAAA}(g, \frac{1}{2}) = \frac{32}{9} - 34.96g + \cdots \\
G^{var}_{AAAA}(g, \frac{1}{2}) = \frac{32}{9} - 31.61g + 368.02g^2 + \cdots \tag{18}
\]

For strong coupling and arbitrary \( c \):

\[
E^{ex}(g, c) = \frac{1}{2} \log g + \frac{1}{2} \log 3 - \frac{3}{4} + \cdots \\
E^{var}(g, c) = \frac{1}{2} \log g + \frac{1}{2} \log 2 + \frac{1}{\sqrt{8g}} + O(\frac{1}{g}) \\
G^{ex}_{AB}(g, c) \to 0 \text{ as } g \to \infty \\
G^{var}_{AB}(g, c) = \frac{c}{2g} - \frac{c}{(2g)^2} + O(\frac{1}{g^2}) \\
G^{ex}_{AAAA}(g, c) = \frac{1}{g} + \cdots \\
G^{var}_{AAAA}(g, c) = \frac{1}{g} - \frac{2}{(2g)^2} + O(\frac{1}{g^2}) \tag{19}
\]
We see that both for strong and weak coupling, our wignerian variational ansatz provides good estimates for the partition function and correlations. We now consider the gaussian + Yang Mills two matrix model, where we compare our variational estimates with Monte-Carlo measurements.

6 Gaussian + Yang-Mills two matrix model

Let us specialize to the two matrix model with action \((m^2 > 0)\)

\[
S(A) = \text{tr} \left[ \frac{m^2}{2} (A_1^2 + A_2^2) - \frac{1}{4} [A_1, A_2]^2 \right]
\]  

(20)

For the wignerian ansatz,

\[
\Omega[G] = \frac{1}{2} \log \det[G_{ij}] - \frac{m^2}{2} (G_{11} + G_{22}) + \frac{1}{2} (G_{1212} - G_{1221})
\]

(21)

Due to the \(A_1 \leftrightarrow A_2\) symmetry of the action we can assume a variational matrix

\[
G_{ij} = \begin{pmatrix} \alpha & \beta \\ \beta & \alpha \end{pmatrix}
\]

(22)

Since \(< \frac{\mu}{N} A_i^2 > \geq 0\) and \(< \frac{\mu}{N} (A_1 - A_2)^2 > \geq 0\), we must maximize

\[
\Omega(\alpha, \beta) = \frac{1}{2} \log(\alpha^2 - \beta^2) - m^2 \alpha + \frac{1}{2} (\beta^2 - \alpha^2)
\]

(23)

in the region \(\alpha \geq 0\) and \(\alpha \geq \beta\). We get

\[
G_{11} = G_{22} = \alpha = \sqrt{1 + \frac{m^4}{4} - \frac{m^2}{2}}, \quad G_{12} = G_{21} = \beta = 0
\]

(24)

Figures 1 and 2 compare the variational two point correlations with Monte-Carlo measurements for a range of values of \(m^2\).

All other correlations can be expressed in terms of these. For example, the 4-point correlations are (the rest are determined by cyclic symmetry and \(A_1 \leftrightarrow A_2\) exchange symmetry)

\[
G_{1111} = 2\alpha^2; \quad G_{1212} = 0; \quad G_{1221} = \alpha^2; \quad G_{1112} = 0
\]

(25)
Figure 1: $\log_{10}[G_{11}]$ versus $\log_{10}[m^2]$. Solid line is variational estimate, dots are the Monte-Carlo measurements. The approximation becomes poor for small values of $m^2$.

Figure 2: $G_{12}$ versus $\log_{10}[m^2]$

Figure 3: $\log_{10}[G_{1111}]$ versus $\log_{10}[m^2]$

Figure 4: $G_{1212}$ versus $\log_{10}[m^2]$
Figures 3, 4 and 5 compare variational estimates (solid lines) and Monte-Carlo measurements (dots) of $G_{1111}$, $G_{1212}$ and $G_{1122}$ for $10^{-3} \leq m^2 \leq 10^3$. The $n$ point pure $A_1$ (or $A_2$) correlation is given by the Catalan numbers

$$G_{11\ldots1} = G_{22\ldots2} \equiv G^{(n)} \equiv \begin{cases} \frac{c_n \alpha^{\frac{n}{2}}}{(\frac{n}{2})!(\frac{n}{2}+1)!} \alpha^{\frac{n}{2}}, & \text{if } n \text{ is even;} \\ 0, & \text{if } n \text{ is odd.} \end{cases}$$

More generally,

$$G^{(n_1\ldots2)} = G^{(n_2\ldots2)} \equiv G^{(n_1)} G^{(n_2)}$$

$$G^{(n_1\ldots2\ldots2)} \equiv G^{(n_1)(n_2)} = G^{(n_1)} G^{(n_2)}$$

$$G^{(n_1\ldots2\ldots2\ldots2)} \equiv G^{(n_1)(n_2)} G^{(n_3)} G^{(n_4)} = G^{(n_1+n_3)} G^{(n_2)} G^{(n_4)}$$

We mention these since they are useful in estimating expectation values of Wilson loop-like operators. The variational estimate for vacuum energy is

$$E_{\text{var}}(g) = -\lim_{N \to \infty} \frac{1}{N^2} \log \left[ \frac{Z(g)}{Z(0)} \right] = -\log \alpha = -\log \left[ \sqrt{1 + \frac{m^2}{4} - \frac{m^2}{2}} \right]$$

6.1 Wilson Loop Operators

It is also interesting to see what the wignerian ansatz says about the 2-matrix analogue of the expectation of the Wilson loop in the large $N$ limit.

**Wilson Line:** The simplest analogue is a ‘Wilson line’ the analogue of the parallel transport along a line of length $l$ in the $A_1$ direction

$$W_{\text{line}}(l) = \lim_{N \to \infty} \frac{1}{N} \langle \text{tr} e^{ilA_1} \rangle$$
Figure 6: $W_{\text{line}}(l)$ for $m^2 = 1$. Dots are numerical and solid line variational estimate.

For the wignerian ansatz, we get (using eqn (26), $J_n(z)$ is the Bessel function of the first kind)

$$W_{\text{line}}(l) = \sum_{n=0}^{\infty} \frac{(il)^{2k}}{(2k)!} c_k \alpha^k$$

$$= \frac{1}{l\sqrt{\alpha}} J_1(2l\sqrt{\alpha}) \sim \frac{1}{\sqrt{\pi l \sqrt{\alpha}}} \cos \left(\frac{3\pi}{4} - 2l\sqrt{\alpha}\right)$$

as $l \to \infty$ (30)

$W_{\text{line}}(l)$ is a real-valued function of real $l$ since the odd order correlations vanish. Thus, for the wignerian ansatz, the expectation value of the ‘Wilson line’ is oscillatory but decays as a power $l^{-3/2}$. For small $l$, $W_{\text{line}}(l) \to 1 - \frac{1}{2} i l^2 + \frac{\alpha^2 l^4}{12} - \cdots$. Figure 6 compares this ansatz with Monte-Carlo measurements for $m^2 = 1$. The behavior both for small and large values of $l$ is well captured by our ansatz.

**L shaped Wilson Line:** For an $L$ shaped curve, we define

$$W_L(l) = \lim_{N \to \infty} \frac{1}{N} \langle tr N e^{il A_1} e^{il A_2} \rangle$$

(31)

For the wignerian ansatz (use eq. 27), $1 F_2(a, b; z) = \sum_{n=0}^{\infty} \frac{(a)_n}{(b)_n} \frac{z^n}{n!}$ is a generalized Hypergeometric function with $(a)_n$ the Pochhammer symbol,

$$W_L(l) = \sum_{n_1, n_2=0}^{\infty} \frac{(il)^{n_1+n_2}}{n_1! n_2!} \lim_{N \to \infty} \frac{1}{N} \langle tr A_1^{n_1} A_2^{n_2} \rangle$$

$$= \sum_{k_1, k_2=0}^{\infty} \frac{(-l^2 \alpha)^{k_1+k_2}}{k_1! (k_1+1)! k_2! (k_2+1)!}$$

$$= \sum_{n=0}^{\infty} \frac{(-l^2 \alpha)^n}{\sqrt{\pi} \Gamma(n+1) \Gamma(n+2) \Gamma(n+3)} \frac{4^{n+1} \Gamma(n+\frac{3}{2})}{\Gamma(n+3)}$$

15
\[ W_L(l) = \, _1F_2\left(\frac{3}{2}; \{2, 3\}; -4l^2\alpha\right) \quad (32) \]

As before \( W_L(l) \) is real for real \( l \). For small \( l \), \( W_L(l) \to 1 - \alpha l^2 + \frac{5\alpha^2 l^4}{2} - \cdots \).

This is compared with the numerical calculation in fig. 7 for \( m^2 = 1 \). Both the small \( l \) behavior and decay for large \( l \) are captured by our estimate.

**Wilson Square:** The analogue of the parallel transport around a square of side \( l \) in the \( A_1 - A_2 \) plane is

\[ W_{\text{square}}(l) = \lim_{N \to \infty} \frac{\text{tr}}{N} e^{ilA_1} e^{ilA_2} e^{-ilA_1} e^{-ilA_2} > \quad (33) \]

In the wignerian variational approximation, \( W_{\text{square}}(l) \) is real-valued since odd order correlations vanish. Using eq.

\[ W_{\text{square}}(l) = \sum_{n=0}^{\infty} (-l^2\alpha)^n T_1(n) + 2(l^2\alpha)^n \sum_{n=0}^{\infty} (-l^2\alpha)^n T_2(n) \quad (34) \]

where,

\[
\begin{align*}
T_1(n) &= \sum_{k_1 \geq 0, k_1 + \cdots + k_4 = n} \frac{2c_{k_1+k_2}c_{k_3+k_4} - \Pi_{i=1}^{4} c_{k_i}}{\Pi_{i=1}^{4}(2k_i)!} \\
T_2(n) &= \sum_{k_1 \geq 0, k_1 + \cdots + k_4 = n} \frac{c_{k_1+k_3+1}c_{k_2+k_4}}{(2k_1+1)(2k_2)(2k_3+1)(2k_4)!} \quad (35)
\end{align*}
\]

For small \( l \), \( W_{\text{square}}(l) \to 1 - l^4\alpha^2 + \frac{5l^6\alpha^3}{6} - \cdots \). The expectation value of the Wilson loop is a rapidly decaying function for large values of \( l \). It would be interesting to find the asymptotic rate of decay. It is oscillatory but a positive function, unlike \( W_{\text{line}} \). These variational predictions are confirmed by the numerics, in fig. 8 for \( m^2 = 1 \).
The variational ansatz does a very good job of estimating the Wilson loop averages over the entire range of values of $l$ studied, for $m^2$ of order unity or more.

7 Summary and Discussion

We find that despite its simplicity, the wignerian ansatz for our entropic variational principle is remarkably good at estimating correlations. For the exactly solved model studied by Mehta, our ansatz works well both for strong and weak coupling. For the gaussian + Yang-Mills two matrix model, our estimates for correlations and expectation values of Wilson loop operators are accurate for moderate and weak coupling when compared with Monte-Carlo measurements. When the commutator squared term dominates the action, the wignerian ansatz becomes poorer as an approximation. This is to be expected. The measured correlations grow without bound as we approach the pure commutator squared limit. This reflects the divergence of the matrix integrals for the pure commutator squared interaction.

We can go beyond the wignerian ansatz for the entropic variational principle. In particular, this will allow us to improve on the wignerian ansatz and also estimate correlations that are identically zero for the wignerian ansatz. We will address this question in a future paper.

In another direction, we hope to extend our approximation methods to supersymmetric matrix models in order to make predictions about the matrix models of M-theory and superstring theory mentioned in the introduction.
A conceptual shortcoming of the numerical procedure used in this paper is that we measured correlations for several finite values of \( N \) before extrapolating to the \( N = \infty \) limit. Is there some way of computationally determining the \( N = \infty \) correlations directly? This is a challenging problem that would likely require new ideas from non-commutative algebra, geometry, probability theory and computer science to determine directly the \( N = \infty \) correlation tensors without integrating over matrix elements.

It would be interesting to make precise the field theoretic connection between general multi-matrix models and supersymmetric gauge theories with adjoint chiral superfields. Our methods for estimating the correlations of the former can potentially shed light on the expectation values of operators in the chiral ring of the supersymmetric gauge theory.

Our investigations have focussed on matrix models with a finite number of matrices. It would be desirable to have a similar theoretical approach based on approximation methods over and above the large \( N \) limit, for matrix field theories (beyond \( c = 1 \) matrix quantum mechanics) such as Yang-Mills theory. This would complement the lattice gauge theory Monte-Carlo efforts (see eg. \cite{39}).

Acknowledgements: The author thanks S.G. Rajeev, A. Agarwal and L. Akant for useful discussions and A. Agarwal for useful comments on the manuscript.

References

[1] M. L. Mehta, *Random Matrices*, 2nd ed. (Academic Press, New York, 1991)

[2] G. ’t Hooft, *Under The Spell Of The Gauge Principle*, Adv. Ser. Math. Phys. **19**, 1 (1994).

[3] E. Witten, Nucl. Phys. B **160**, 57 (1979).

[4] A. A. Migdal, Phys. Rept. **102**, 199 (1983).

[5] A. Jevicki and B. Sakita, Nucl. Phys. B **165**, 511 (1980).

[6] P. Cvitanovic, P. G. Lauwers and P. N. Scharbach, Nucl. Phys. B **203**, 385 (1982).
[7] T. Eguchi and H. Kawai, Phys. Rev. Lett. 48, 1063 (1982).

[8] P. Ginsparg and G. W. Moore. arXiv:hep-th/9304011

[9] D. V. Voiculescu, Bull. London Math. Soc. 34, no. 3, 257–278. (2002)

[10] T. Banks, W. Fischler, S. H. Shenker and L. Susskind, Phys. Rev. D 55, 5112 (1997) arXiv:hep-th/9610043.

[11] N. Ishibashi, H. Kawai, Y. Kitazawa and A. Tsuchiya, Nucl. Phys. B 498, 467 (1997) arXiv:hep-th/9612115.

[12] W. Krauth, H. Nicolai and M. Staudacher, Phys. Lett. B 431, 31 (1998) arXiv:hep-th/9803117.

[13] W. Krauth, J. Plefka and M. Staudacher, Class. Quant. Grav. 17, 1171 (2000) arXiv:hep-th/9911170.

[14] J. Ambjorn, K. N. Anagnostopoulos, W. Bietenholz, T. Hotta and J. Nishimura, JHEP 0007, 011 (2000) arXiv:hep-th/0005147.

[15] J. Nishimura, T. Okubo and F. Sugino, JHEP 0210, 043 (2002) arXiv:hep-th/0205253.

[16] G. M. Cicuta, L. Molinari and G. Vernizzi, J. Phys. A 35, L51 (2002) arXiv:hep-th/0109160.

[17] P. Austing and J. F. Wheater, JHEP 0102, 028 (2001) arXiv:hep-th/0101071.

[18] R. Dijkgraaf and C. Vafa, Nucl. Phys. B 644, 3 (2002) arXiv:hep-th/0206255.

[19] R. Dijkgraaf and C. Vafa, arXiv:hep-th/0208048

[20] F. Cachazo, M. R. Douglas, N. Seiberg and E. Witten, JHEP 0212, 071 (2002) arXiv:hep-th/0211170.

[21] V. A. Kazakov and A. Marshakov, J. Phys. A 36, 3107 (2003) arXiv:hep-th/0211236.
[22] V. A. Kazakov, I. K. Kostov and N. A. Nekrasov, Nucl. Phys. B 557, 413 (1999) [arXiv:hep-th/9810035].

[23] M. R. Douglas, I. R. Klebanov, D. Kutasov, J. Maldacena, E. Martinec and N. Seiberg, [arXiv:hep-th/0307195]

[24] J. McGreevy and H. Verlinde, [arXiv:hep-th/0304224]

[25] A. Sen, [arXiv:hep-th/9904207]

[26] E. Brezin, C. Itzykson, G. Parisi and J. B. Zuber, Commun. Math. Phys. 59, 35 (1978).

[27] M. L. Mehta, Commun. Math. Phys. 79, 327 (1981).

[28] M. Staudacher, Phys. Lett. B 305, 332 (1993) [arXiv:hep-th/9301038].

[29] V. A. Kazakov and P. Zinn-Justin, Nucl. Phys. B 546, 647 (1999) [arXiv:hep-th/9808043].

[30] L. Akant, G. S. Krishnaswami and S. G. Rajeev, Int. J. Mod. Phys. A 17, 2413 (2002) [arXiv:hep-th/0111263].

[31] A. Agarwal, L. Akant, G. S. Krishnaswami and S. G. Rajeev, Int. J. Mod. Phys. A 18, 917 (2003) [arXiv:hep-th/0207200].

[32] A. Agarwal and S. G. Rajeev, Mod. Phys. Lett. A 17, 481 (2002) [arXiv:hep-th/0202095].

[33] A. Agarwal, L. Akant, G. S. Krishnaswami and S. G. Rajeev, in preparation.

[34] N. Metropolis, A. W. Rosenbluth, M. N. Rosenbluth, A. H. Teller and E. Teller, J. Chem. Phys. 21, 1087 (1953).

[35] M. Creutz, Quarks, Gluons And Lattices, Cambridge University Press, Cambridge (1983).

[36] G. Bhanot, Rept. Prog. Phys. 51, 429 (1988).

[37] C. Itzykson and J. M. Drouffe, Statistical Field Theory. Vol. 2 Cambridge University Press (1998).
[38] J. Shao and D. Tu, *The Jackknife and Bootstrap*, Springer Verlag (1995).

[39] B. Lucini and M. Teper, Phys. Rev. D 66, 097502 (2002) [arXiv:hep-lat/0206027].