Diminishing parochialism in intergroup conflict by disrupting the right temporo-parietal junction

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Individuals react to violation of social norms by outgroup members differently than to transgressions of those same norms by ingroup members: namely outgroup perpetrators are punished much more harshly than ingroup perpetrators. This parochial punishment pattern has been observed and extensively studied in social psychology and behavioral economics. Despite progress in recent years, however, little is known about the neural underpinnings of this intergroup bias. Here, we demonstrate by means of transcranial magnetic stimulation (TMS) that the transient disruption of the right, but not the left temporo-parietal junction (TPJ), reduces parochial punishment in a third-party punishment paradigm with real social groups. Moreover, we show that this observed TMS effect on parochial punishment is mediated by a classical punishment motive, i.e., retaliation. Finally, our data suggests that a change in perspective-taking might be the underlying mechanism that explains the impact of right TPJ disruption on retaliation motivation and parochial punishment. These findings provide the first causal evidence that the right TPJ plays a pivotal role in the implementation of parochial behaviors.

Keywords: parochialism; outgroup hostility; ingroup love; normative behavior; third-party punishment; transcranial magnetic stimulation; temporo-parietal junction; intergroup bias

INTRODUCTION
Parochialism, i.e., the differential treatment of ingroup and outgroup members, constitutes a pervasive psychological phenomenon that has a strong impact on human society. It is qualified by a preference for altruistic behavior toward the members of one’s ethnic, racial or any other social group, combined with a tendency for indifference, mistrust or even hostility toward outgroup members (Brewer, 1999; Hewstone & Brown, 2002). For example, people are more likely to help ingroup than outgroup members when they are in an emergency situation (Levine et al., 2005) or when they can relieve these in- and outgroup members from pain (Hein et al., 2010); they cooperate more with ingroup than with outgroup members (Tajfel et al., 1971; Brewer, 1979; Goette et al., 2006; Fehr et al., 2008; Ben-Ner et al., 2009; Chen and Li, 2009; Koopmans and Rebers, 2009), and they punish norm violations less strongly when ingroup members rather than outgroup members commit them (Bernhard et al., 2006; Goette et al., 2006). Hence, the differential behaviors toward members of different social groups underline acts such as discrimination, terrorism or open intergroup conflict (Bowles, 2009; Bruno and Saxe, 2010; De Dreu et al., 2010; Haushofer et al., 2010).

Recent evidence from functional magnetic resonance imaging (fMRI) studies suggests that areas involved in social cognition (Gallagher and Frith, 2003; Lieberman, 2007; Van Overwalle, 2009), including the dorsomedial prefrontal cortex (DMPFC) and bilateral temporo-parietal junction (TPJ), might play a role in differentiating between ingroup and outgroup members in perception (Harris and Fiske, 2006), judgments (Freeman et al., 2010; Mathur et al., 2010; Falk et al., 2012) and behavior (Baumgartner et al., 2012).

However, functional imaging studies, although indispensable, do not permit causal inferences about the effect of brain processes on human perception, judgment and behavior, because the observed neural activations could simply be an epiphenomenon or a consequence, and not necessarily the cause of the perception, judgment or behavior. In contrast, brain stimulation techniques, such as transcranial magnetic stimulation (TMS), which interfere non-invasively with the activity of defined areas in the human cortex, allow researchers to draw causal conclusions about the behavioral impact of the stimulated brain region. No previous study, however, has used brain stimulation techniques to provide causal evidence about the impact of brain areas on parochialism. Using inhibitory low-frequency repetitive TMS (rTMS), we disrupted two key areas involved in social cognition (i.e., right and left TPJ) in order to investigate the causal role these areas play in the differential behavior toward outgroup and ingroup members.

For that purpose, we applied a third-party punishment paradigm with real social groups (soccer fans, see ‘Materials and Methods’ section for details) and real monetary stakes and consequences for all interaction partners. Third-party punishment is argued to be one of the decisive factors regulating the enforcement of social norms in human society (Bendor and Swistak, 2001; Fehr and Fischbacher, 2004) because third-parties are willing to punish norm violators despite the fact that the norm transgression does not directly affect them. Crucially for this study, third-party punishment has been shown to be strongly shaped by parochialism (Bernhard et al., 2006; Goette et al., 2006; Baumgartner et al., 2012), and lacks any strategic component which could otherwise obscure the parochial behavioral tendencies.

The applied third-party punishment paradigm consists of two decision stages. The first stage took place in the form of an Internet experiment a few weeks before the second stage. Soccer fans were paired with either a member of their own team (ingroup) or someone from a rival team (outgroup) and they played a simultaneous Prisoner’s Dilemma Game (PDG) in the roles of either Player A or B. Each player was endowed with 20 points and each had to decide simultaneously whether to keep the points or pass all of them to the other player. Passed points were doubled. Thus, keeping the points equals defection (denoted as D throughout the article) and passing the points equals cooperation (denoted as C throughout the article). For example, if A retained the 20 points while B transferred the 20 points (behavioral pattern DC), then A earned a total of 60 points (40 points from the transfer plus the original 20 points) and B earned nothing.
In the second stage, which took place in the TMS laboratory, subjects in the role of a third-party (Player C) received the opportunity to punish Player A’s or Player B’s behavior by assigning punishment points. For that purpose, Player C received an endowment of 10 points at the beginning of each trial, which C could use to finance the assignment of punishment points. Assigning one punishment point costs Player C one point and the sanctioned player three points. Importantly, Player C only could punish the behavior of one player (either A or B) during each of the punishment trials played. In order to simplify the nomenclature, we recoded all Player C’s decisions in such a way that A always refers to the player that C can punish, while B always refers to the player that C cannot punish. Player C could retain any points not used for punishment as income (Figure 1).

As we wanted to examine the causal impact of the right and left TPJ on third-parties’ parochialism, third-parties were confronted with different combinations of Player A’s and B’s group affiliations: (i) Both Player A and Player B were ingroup members (referred to in the following as IN/IN), (ii) Player A was an ingroup member and Player B was an outgroup member (referred to as IN/OUT) or (iii) Player A was an outgroup member and Player B an ingroup member (referred to as OUT/IN) (please see Figure 1, and also the supplementary Figure 1 for an example of a decision screen).

With this design, we were able to explore the impact of rTMS on parochial punishment patterns in an intergroup constellation and in an intragroup constellation (Figure 1). The most pronounced and explored parochial punishment pattern in the literature is the stronger punishment of an outgroup member who defects against an ingroup member compared with the less harsh punishment of an ingroup member who defects against an outgroup member (difference value OUT/IN − IN/OUT). This punishment pattern represents the comparison of Player C’s level of punishment against outgroup perpetrators with that against ingroup perpetrators. We therefore label this punishment pattern as parochial punishment in an ‘intergroup constellation’ (Figure 1). A few studies (e.g. Bernhard et al., 2006) further reported stronger punishment of an ingroup member who defects against an ingroup member compared with an ingroup member who defects against an outgroup member (difference value IN/IN − IN/OUT). Both perpetrators in this punishment pattern stem from the same social group, i.e. ingroup. We therefore label this ‘intragroup constellation’. Thus, by creating these difference values, our study allows us to examine whether parochial punishment behavior is causally modulated by rTMS of the right or left TPJ with regard to the intergroup constellation, the intragroup constellation or both. We expected a potential rTMS effect to be most pronounced in the most conflicting trials, i.e. in DC trials (Player A, who can be punished, defects and Player B cooperates), because studies consistently indicate that this is the situation where the strongest parochial punishment occurs (Goette et al., 2006; Baumgartner et al., 2012).

In order to answer these research questions, we applied low-frequency rTMS for 20 min to either the right TPJ or the left TPJ or sham TMS to healthy subjects in the role of Player C. Low-frequency rTMS for the duration of several minutes leads to a suppression of activity in the stimulated brain regions that outlasts the duration of the TMS train for about half of the duration of the stimulation (Eisenegger et al., 2008; Baumgartner et al., 2011). Each subject participated in only one of the three treatment conditions, and none had experienced TMS previously. We used a stereotaxic infrared system (Brainsight) to localize the right or left TPJ on each subject’s individual anatomical brain scan based on the peak coordinates reported in our functional neuroimaging study on parochialism (Baumgartner et al., 2012).

Taken together, our study allows us to clarify the following research questions: First, are the right TPJ, left TPJ or both areas causally involved in parochialism? Second, if we find such causal evidence, what is the direction of the effect (i.e. a more pronounced parochial punishment pattern or less differential behavior toward ingroup and outgroup members)? Third, does rTMS affect all parochial punishment patterns (i.e. parochialism in the intergroup constellation and intragroup constellation) similarly? Fourth, we measured several emotions and motivations (anger, justice, improvement of future behavior and retaliation) known to affect punishment decisions (Keller et al., 2010). Thus, if we observe an impact of rTMS on one or both of the
parochial punishment patterns, are we able to identify a punishment-related emotion or motivation that mediates the effects of disrupting the right or left TPJ on parochialism?

**MATERIALS AND METHODS**

**Subjects**
We studied 36 healthy men (mean age ± s.d. = 24.3 ± 4.2 years). Subjects gave informed written consent prior to participating in the study, which was approved by the local ethics committee. Subjects were randomly assigned to one of the three treatment conditions (rTMS of left TPJ = 12 subjects, rTMS of right TPJ = 13 subjects and sham stimulation = 11 subjects), and none had experienced TMS or a third-party punishment game previously. No subject had a history of psychiatric illness or neurological disorders. Subjects neither experienced serious adverse side effects nor reported scalp pain, neck pain or headaches after the experiment.

**Transcranial magnetic stimulation**
We applied low-frequency (1 Hz) rTMS to the TPJ for 20 min (1200 pulses) before subjects participated in the third-party punishment paradigm (off-line paradigm), using a Magstim Rapid Magnetic Stimulator (Magstim, Winchester, MA, USA) and a commercially available figure-of-eight coil (70-mm diameter double-circle, air-cooled). To investigate a possible hemispheric laterality in the role of the TPJ on parochialism, we applied rTMS to the right TPJ or to the left TPJ. The creation of stimulation groups receiving rTMS to the same homologous area, but on different hemispheres, is important to control for the potential side effects of rTMS (Abler et al., 2005), including discomfort, irritation and mood changes (see Supplementary Table 2 and Supplementary Figure 5 for statistical evidence that this control procedure was in fact successful). Furthermore, we had a control condition in which we applied sham stimulation for 20 min to the right or left TPJ.

In order to localize the right or left TPJ, we acquired individual anatomical brain images (T1-weighted) of each subject. These individual brain images were used to plan, guide and monitor the stimulation in real time using a stereotactic infrared system (Brainsight), ensuring that every pulse was delivered to the predetermined cortical location (Sack et al., 2009). The coordinates for stimulation were taken from our functional neuroimaging study (Baumgartner et al., 2012), in which we found correlational evidence that the areas in the left and right TPJ might play a role in parochialism (MNI-coordinates for left TPJ: $x = -45, y = -60, z = 21$, MNI-coordinates for right TPJ: $x = 57, y = -60, z = 30$). We transformed these MNI-coordinates to each subject’s native brain space using the Brainsight software, in order to make them suitable for neuronavigation with the stereotactic infrared system.

Stimulation intensity was set at 110% of the individual resting motor threshold (MT), as determined following current guidelines (Rossi et al., 2009). We did not observe any differences with respect to stimulation intensity for the left vs right TMS groups (independent $t$-test: $t_{(23)} = -0.385, P = 0.70$). For the left TPJ, stimulation intensity was on average at 61.6% of maximal stimulator output, and for the right TPJ, the stimulation intensity was on average at 62.3% of maximal stimulator output. The coil was held tangentially to the subject’s head with the handle pointing caudally. The rTMS parameters are well within currently recommended guidelines (Rossi et al., 2009) and result in a suppression of excitability of the targeted cortical region for several minutes following completion of the rTMS train (Eisenegger et al., 2008; Baumgartner et al., 2011). The whole paradigm lasted ~9 min (including the questions regarding emotions, motives and perspective-taking) and was thus well within the borders of the rTMS after-effect.

**Social groups and ingroup identification scale**
To explore the impact of rTMS on parochialism, we decided to use naturally occurring social groups. For that purpose, we recruited strong fans of two soccer teams. These two soccer teams are among the best teams in Switzerland (they have won most of the recent championships) and there is a strong rivalry between the two fan groups. Recent work showed that sports team loyalty is one of the most powerful sources for group discrimination and parochialism (Ben-Ner et al., 2009). In order to measure each subject’s strength of identification with their favored soccer team, we applied the Sport Spectator Identification Scale (5-point Likert scale; Wann and Branscombe, 1993). The analysis of this identification scale revealed that all recruited subjects in fact strongly identified with their favored soccer team (mean ± s.e.m. = 3.64 ± 0.10) and that there were no significant differences between the three treatment groups (left TMS, right TMS and sham, for details please see ‘Results’ section).

**The third-party punishment paradigm**
We applied a third-party punishment paradigm with real social groups and the involvement of real monetary stakes. The paradigm consisted of two decision stages, one conducted on the Internet using an online questionnaire tool (Unipark) and one conducted in the TMS laboratory. The subjects knew in all stages of the paradigm that there were no repeated interactions and that all interactions were conducted in complete anonymity in order to exclude reputation effects. In the first stage, subjects in the role of Player A or B played a simultaneous PDG with ingroup and outgroup members, where they could decide either to cooperate or defect. Thus, four behavioral patterns were possible: Players A and B cooperate (CC), Players A and B defect (DD), Player A cooperates and Player B defects (CD), and Player A defects and Player B cooperates (DC). In the second stage (in the TMS laboratory), subjects in the role of Player C were confronted with 30 of these decisions and had to decide whether to punish Player A’s behavior by assigning punishment points. For that purpose, they received an endowment of 10 points at the beginning of each punishment trial which they could either keep or use to punish Player A. One point assigned for punishment reduced the punished player’s income by three points. Points not used for punishment were exchanged into real money and paid to Player C at the end of the experiment (10 points = 1 Swiss Franc, ~$1). Player C was always informed about the group affiliations of Players A and B, that is whether the players were from his own or another social group. The PDG decisions of Players A and B were prespecified such that each Player C was confronted with the same 30 decision situations, which were presented in random orders (for details, please see Supplementary Table 3). The group affiliation and the behavioral decisions of Players A and B were presented both verbally and visually. Please see Supplementary Figure 1 for an example of a decision screen third-parties saw during the TMS session. The software package z-Tree (Fischbacher, 2007) was used for presenting these screens and for collecting behavioral and timing data.

Finally, please note that all Players C took also part in the PDG. The PDG session took place 6–8 weeks before the TMS session was conducted. At the point in time when Players C played the PDG, they did not know anything about the later TMS session. The advantage of this procedure (i.e. Players C experiencing both situations) is 2-fold: first, the experience of the online PDG substantially increases the credibility of the third-party session conducted in the TMS laboratory where Players C were confronted with real decisions of other players made in the online PDG. Second, it enabled us to investigate pre-existing differences in parochial tendencies across treatment groups.
In case of significant multivariate effects, statistical tests conducted, including independent, dependent and one sample t-tests, chi-square tests as well as repeated measures ANOVAs. Results were considered significant at the level of \( P < 0.05 \) (two-tailed). In case of significant multivariate effects, post hoc paired t-tests were computed using the Bonferroni correction according to Holm (1979). As effect size measure \( \eta^2 \) is reported. We controlled for different covariates in our analyses, including trait questionnaires, ingroup identification score and payment in the PDG. Crucially, all reported findings hold, irrespective of whether we controlled for the mentioned covariates (see ‘Results’ section for details).

The SPSS macro ‘Process’ (Preacher and Hayes, 2008; Hayes, 2013) was used for all mediation analyses. It investigates whether an independent variable (\( X \), in our case, the treatment group variable) affects a dependent variable (\( Y \), in our case, the parochial punishment pattern) through one or more intervening variables, or mediators (\( M \), in our case, the item on perspective-taking and/or the punishment-related emotions/motivations). Significance of the mediated, indirect effect through \( M \) was tested with a bootstrapping strategy (Preacher and Hayes, 2008). Bootstrapping involves the repeated extraction of samples from the data set and the estimation of the indirect effect in each resampled data set. By repeating this process thousands of times, an empirical approximation of the sampling distribution of the indirect effect is built and used to construct confidence intervals. We used 10,000 bootstrap samples to generate bootstrap confidence intervals (90%, 95% and 99%) for the indirect effects.

### Results

#### Parochial punishment patterns

In clear agreement with the literature, we found on average, irrespective of treatment groups, a strong impact of the perpetrator’s group membership on third-parties’ invested punishment points in the intergroup constellation, i.e. a much more harsh punishment of an outgroup perpetrator in OUT/IN compared with an ingroup perpetrator in IN/OUT, which was most pronounced in the DC trials (mean punishment differences DC trials \( \pm \text{s.e.m.} = 4.55 \pm 0.55 \), dependent t-test: \( t_{(35)} = 8.1, P < 0.001 \)), and less pronounced in the DD trials (mean punishment difference DD trials \( \pm \text{s.e.m.} = 2.48 \pm 0.50 \), dependent t-test: \( t_{(35)} = 4.9, P < 0.001 \)). Thus, findings indicate that an outgroup member who defects against a cooperating (or defecting) ingroup member is much more severely punished than an ingroup member who defects against a cooperating (or defecting) outgroup member. Furthermore, we found a similar parochial punishment pattern on average, irrespective of treatment groups, in the intragroup constellation, albeit as expected at a lower level and only in DC trials, i.e. an ingroup perpetrator in IN/IN was more strongly punished than an ingroup perpetrator in IN/OUT (mean punishment difference DC trials \( \pm \text{s.e.m.} = 0.84 \pm 0.28 \), dependent t-test: \( t_{(35)} = 2.93, P = 0.006 \)). Thus, an ingroup member who defects against a cooperating ingroup member is punished much more strongly than an ingroup member who defects against a cooperating outgroup member.

Of primary interest for this study is the question whether rTMS of right and/or left TPJ might be able to affect these parochial punishment patterns. Accordingly, we calculated a repeated-measures ANOVA with the within-subject factor behavioral trials (CC, CD, DC and DD) and the between-subject factor treatment group (left TMS, right TMS and sham) for the two parochial punishment patterns (OUT/IN–IN/OUT and IN/IN–OUT/OUT).

Results provide evidence that rTMS in fact modulates the stronger punishment of an outgroup perpetrator in OUT/IN compared with an ingroup perpetrator in IN/OUT, indicated by a significant main effect of treatment group \( (F_{(2,32)} = 3.38, P = 0.046, \eta^2 = 0.17) \) and a significant interaction effect of treatment group × behavioral trials \( (F_{(6,96)} = 2.65, P = 0.020, \eta^2 = 0.14) \). Post hoc independent t-tests revealed that the right TMS group exhibits significantly diminished parochial punishment behavior than either the left TMS group \( (P = 0.009) \) or the sham group \( (P = 0.041) \). This was unique to the DC trials (Figure 2A). A similar pattern emerged in the DD trials, but only the difference between the right TMS and left TMS group

#### Trait questionnaires

In order to corroborate our TMS findings (see ‘Results’ section for details), we measured the following well-established trait questionnaires. The moral foundations questionnaire (Graham et al., 2011) is a 20-item questionnaire (short version), which was developed on the basis of a theoretical model of five universally available (but variably developed) sets of moral intuitions: Harm/Care, Fairness/Reciprocity, Ingroup/Loyalty, Authority/Respect and Purity/Sanctity. It measures on six-point Likert scales how strongly a subject weights these different moral intuitions when deciding whether something is right or wrong. The social dominance orientation questionnaire (Jost and Thompson, 2000) is a two-dimensional 16-item approach to the measurement of social dominance orientation using 7-point Likert scales. One dimension measures general opposition to equality and the other dimension measures support for group-based dominance. We also applied six items of the personal norm of reciprocity scale (Perugini et al., 2003) which measures (on two 7-point Likert scales) a subject’s tendency to reward another person’s positive behavior (positive reciprocity, three items) and to punish negative behavior (negative reciprocity, three items). Finally, we applied the injustice sensitivity scale from the observer perspective (Schmitt et al., 2005). This 10-item questionnaire measures on a 7-point Likert scale how strongly subjects in the role of an observer (i.e. a third-party) psychologically and emotionally react to observed injustice. All trait questionnaires were measured 30 min after the end of the TMS administration, i.e. when the rTMS effect was expected to have dissipated (Eisenegger et al., 2008).

#### Punishment-related emotions and motivations

In order to examine whether disrupting key areas of social cognition impacts punishment-related emotions and motivations, we obtained ratings for four emotions and motivations known to play a role in punishment-related decision-making (anger, retaliation, improvement of future behavior and justice; Keller et al., 2010). At the end of the punishment paradigm, subjects were confronted with some of the previously experienced decision situations and were asked to indicate on 7-point Likert scales how strongly these emotions and motivations had affected their decision to assign punishment points. Due to time limits of the TMS after-effect, we only obtained these emotions/motives for DC trials (Player A, who can be punished, defects, whereas Player B cooperates). We decided to focus on these trials, because here the strongest parochial punishment patterns can be found (Bernhard et al., 2006; Baumgartner et al., 2012).

#### Item on perspective-taking

After the ratings of the punishment-related emotions and motivations, subjects were asked to answer the following statement regarding trials with unilateral defection (DC trials) committed by an ingroup or an outgroup perpetrator in the role of Player A: ‘Putting myself in the position of Player A helped me to make my punishment decision’. Subjects had to indicate on 6-point Likert scales whether they agreed with the statement or not (ranging from ‘I agree not at all’ to ‘I agree completely’).

#### Statistical analyses

All statistical analyses were run with SPSS 20 for PC (SPSS Inc., Chicago, IL, USA). Please see ‘Results’ section for details about the statistical tests conducted, including independent, dependent and one sample t-tests, chi-square tests as well as repeated measures ANOVAs. Results were considered significant at the level of \( P < 0.05 \) (two-tailed). In case of significant multivariate effects, post hoc paired t-tests were computed using the Bonferroni correction according to Holm (1979). As effect size measure \( \eta^2 \) is reported. We controlled for different
was significant ($P = 0.011$), while the difference between the right TMS and sham group did not reach statistical significance ($P = 0.260$). A closer inspection of the observed TMS effect revealed that the diminished parochial punishment is mainly caused by a reduction in OUT/IN, i.e. third-parties with disrupted right TPJ show a less severe punishment of an outgroup perpetrator who defects against a cooperating ingroup member (see Supplementary Figure 2A). Finally, as expected, there were no TMS effects with regard to the CC and CD trials, where punishment was virtually absent in all treatment groups (all $P > 0.18$).

Next, we checked whether we find evidence in our data that rTMS causally modulates the stronger punishment of an ingroup perpetrator in IN/IN compared with an ingroup perpetrator in IN/OUT. Results of the repeated-measures ANOVA revealed neither a main effect of treatment group ($F_{(2,33)} = 1.553$, $P = 0.23$, $ETA^2 = 0.08$) nor an interaction effect of treatment group x behavioral trials ($F_{(6,99)} = 1.473$, $P = 0.21$, $ETA^2 = 0.08$), suggesting that the stronger punishment in the intragroup constellation, i.e. the stronger punishment of an ingroup perpetrator in IN/IN compared with an ingroup perpetrator in IN/OUT, is neither causally modulated by rTMS of the right nor the left TPJ.

**Punishment-related emotions and motivations as mediators**

We next investigated whether punishment-related emotions or motivations (anger, retaliation, improvement of future normative behavior and justice) might provide an explanation for the parochial punishment patterns and whether they might mediate the observed TMS effect of disrupting the right TPJ on parochialism. Due to time limitations of the TMS after-effect, we only obtained these emotions and motives for the DC trials, where we expected and found the strongest parochial punishment patterns. Interestingly, only one of the four obtained emotions/motives, namely retaliation, shows a treatment group effect and is capable of mediating the observed TMS effect.

Third-parties report a more severe retaliation motive in OUT/IN vs IN/OUT (dependent $t$-test: $t_{(35)} = 6.90$, $P < 0.001$), i.e. the motive to retaliate against an outgroup perpetrator who defected against an ingroup member is much stronger compared with an ingroup member who defected against an outgroup member. In contrast, we find no difference with regard to this motive in the intragroup constellation (i.e. IN/IN vs IN/OUT; dependent $t$-test: $t_{(35)} = 0.00$, $P = 1.00$). Moreover, we found a clear TMS effect on the retaliation motive (but not on the other emotion/motivations; univariate ANOVA with between-subject factor treatment group: $F_{(2,33)} = 4.00$, $P = 0.028$, $ETA^2 = 0.20$, all other univariate ANOVAs regarding the other emotion/motives: $P > 0.10$). In other words, the perpetrator’s group membership has a stronger impact on the retaliation motive in both the left TMS and sham group compared with the right TMS group (independent $t$-tests: left TMS vs right TMS: $P = 0.01$, sham vs right TMS: $P = 0.05$, Figure 2B). This means that the disruption of the right TPJ by TMS seems to narrow the difference in motivation to retaliate against out- and ingroup perpetrators. A closer inspection of this TMS effect revealed that an attenuated retaliation motive in OUT/IN mainly causes this narrowed difference, i.e. third-parties whose right TPJ was disrupted reported a reduced retaliation motive toward an outgroup perpetrator who defected against a cooperating ingroup member compared with the sham group or left TPJ group (see Supplementary Figure 2B).

In order to test whether the retaliation motive mediates the TMS effect on parochial punishment behavior, we conducted a mediation analysis (Preacher and Hayes, 2008). For this analysis, we merged the left TMS and sham groups into one group, because as shown above, these two groups demonstrated both an undistinguishable parochial punishment pattern and retaliation motive. The path diagram of the mediation analysis (Figure 3) shows that disrupting the right TPJ reduces the difference in the retaliation motive toward an outgroup compared with an ingroup perpetrator which in turn causes a smaller difference in the punishment meted out to an outgroup or ingroup perpetrator in DC trials. Please note that if we conduct two separate analyses for the left TMS and sham group, both mediation analyses are also significant (see Supplementary Figure 3 for details). Moreover, if we conduct mediation analyses with the other emotions/motives as mediators (anger, justice and improvement of future behavior), none of the mediation analyses becomes significant (bootstrapping statistics comparing Path $c'$ vs Path $c$; all $P > 0.20$).

**Possible process that might explain the observed TMS effect on retaliation motive and punishment behavior**

In order to investigate potential functional processes that might be affected by the disruption of the TPJ, we obtained an item on
Fig. 3. Mediation analysis. Depicted is the path diagram (including regression coefficients ± s.e.m. and P-values) of the mediation analysis demonstrating that the retaliation motive mediates the impact of disrupting the right TPJ on parochial punishment pattern in a conflicting intergroup constellation. All four requirements for a mediation effect are satisfied: Path a, Path b, and Path c are significant, and Path c’ is significantly smaller than Path c. In detail, Path a represents the effect of TMS (right TPJ) vs sham on the retaliation motive. Path b represents the impact of the retaliation motive on parochial punishment controlling for the TMS effect. Together, Path a and Path b represent the indirect (mediated) effect of TMS on parochial punishment through the retaliation motive. Path c represents the direct effect of TMS on parochial punishment and is calculated controlling for the indirect, mediated effect. Path c’ represents the total (mediated and direct) effect of TMS on parochial punishment. Finally, the decisive statistical test to examine whether mediation occurs is the statistical test between Path c and Path c’. Bootstrapping statistics (see ‘Materials and Methods’ section for details) revealed that Path c’ is significantly smaller than Path c (P < 0.01), providing evidence that the retaliation motive is indeed a significant mediator.

No treatment group differences in dispositional traits, parochial patterns in the PDG and ingroup identification scores

In a final set of analyses, we checked whether we can provide evidence that the observed treatment group effect is not due to pre-existing differences in dispositional characteristics across treatment groups. We measured a series of trait questionnaires (at a time when the TMS effect can be expected to have worn off) with regard to the sensitivity toward injustice (injustic sensitivity scale; Schmitt et al., 2003), propensity to reciprocate kind or hostile acts (personal norm of reciprocity, Perugini et al., 2003), moral foundations (moral foundations questionnaire, Graham et al., 2011) and social dominance orientation (Jost and Thompson, 2000). Univariate ANOVAs revealed no difference in any scale across treatment groups (all P > 0.11, please see Supplementary Table 1 for details). Furthermore, if we use all these scales as covariates in the two reported analyses showing treatment group effects in DC trials, both treatment group effects remain significant (parochial punishment in intergroup constellation: F(2, 23) = 3.92, P = 0.035, ETA^2 = 0.26; retaliation motive in intergroup constellation: F(2, 23) = 5.95, P = 0.01, ETA^2 = 0.34). A few weeks before subjects took part in the TMS session, they also played the described PDG. Since this paradigm also evokes strong parochial patterns (i.e. stronger cooperation rate with ingroup members, see for example Goette et al., 2006; Koopmans and Rebers, 2009), we have the possibility to check whether there were pre-existing differences with regard to parochial behavior across treatment groups. Findings revealed that all treatment groups demonstrated a similar parochial behavior, qualified by a much higher cooperation rate when interacting with ingroup members compared with outgroup members (difference in cooperation rate, i.e. ingroup > outgroup: left TMS group: 75%, right TMS group: 70% and sham group: 73%). Chi-square tests revealed no difference across treatment groups regarding the cooperation rates in the PDG (all P > 0.89). In other words, we found no pre-existing differences across treatment groups in the PDG. Finally, because studies on ingroup and outgroup evaluations showed that ingroup biases in judgments depend on how strongly subjects identify with ingroup members (Aberson et al., 2000), we also checked for potential pre-existing differences with regard to ingroup identification. We used an ingroup identification questionnaire (Wann and Branscombe, 1993; 5-point Likert scale) and found no differences across treatment groups (univariate ANOVA with between-subject factor treatment group: F(2, 23) = 0.15, P = 0.86, ETA^2 = 0.01). Moreover, if we use this identification score as covariate, the observed TMS effects on parochial punishment (repeated measures ANOVA: F(2, 31) = 3.69, P = 0.036, ETA^2 = 0.19) and retaliation motive (repeated measures ANOVA: F(2, 32) = 4.19, P = 0.024, ETA^2 = 0.21) remain significant.

Taken together, these findings provide strong evidence that the observed TMS effect can be explained neither by pre-existing differences in dispositional characteristics nor by the strength of ingroup identification.
DISCUSSION

Our results indicate a lateralized effect of a disrupted function of the TPJ on parochialism in intergroup conflict. We found that disrupting the function of the right, but not left TPJ, significantly diminishes the impact of the perpetrator’s group membership on punishment decisions. Third-parties whose right TPJ was disrupted by rTMS treat an outgroup perpetrator who defected against a cooperating ingroup member and an ingroup perpetrator who committed the very same norm violation toward an outgroup member more similarly than do third-parties with a disrupted left TPJ or sham stimulation. Our findings further show that third-parties’ punishment behavior in the context of an intragroup constellation is not affected by TMS of the right or left TPJ. Thus, the findings of our study suggest that the TPJ plays a decisive role in modulation of parochialism in intergroup conflict, but not in intragroup conflict.

Furthermore, we find that the retaliation motive mediates the impact of the disrupted right TPJ on parochial punishment. Disrupting the function of the right TPJ seems to narrow the difference in motivation to retaliate against outgroup and ingroup perpetrators, which in turn reduces the parochial punishment magnitude. Finally, we find that disruption of the right TPJ induced a change in mentalizing processes, i.e. the right TMS group reported no differences in regard to how well they were able to put themselves into the position of an ingroup and outgroup perpetrator, i.e. they process the intentions and goals behind the defective act more similarly. As a consequence of this reduced mentalizing bias, the strong retaliation motive toward an outgroup perpetrator is attenuated, which in turn reduces the punishment magnitude.

This study is based on the results of a previous functional neuroimaging study on parochialism (Baumgartner et al., 2012), in which we found increased activity in the left TPJ, right TPJ and DMPFC when third-parties were confronted with defecting ingroup members compared with defecting outgroup members. This study was set up to draw causal conclusions about the necessity of the TPJ in parochial punishment. Our findings suggest that this is the case for the processes in the right, but not left TPJ. The fact that no TMS effect is observed after disruption of the left TPJ indicates that this area does not play a causal role in parochial punishment. Since our previous fMRI study on parochial punishment linked the functional connectivity between left TPJ and DMPFC to the punishment of ingroup perpetrators (i.e. the stronger the connectivity, the lower the punishment of ingroup perpetrator), we might speculate that the DMPFC alone is ‘capable’ of implementing the processes required for the punishment of ingroup perpetrators and does not require the additional contribution of the left TPJ. Future research could investigate how disrupting the function of the DMPFC impacts parochial punishment behavior.

The diminished parochialism in subjects with a disrupted right TPJ was evident in the fact that outgroup perpetrators were treated more like ingroup perpetrators, i.e. punished less strongly and not due to treating ingroup perpetrators more like outgroup perpetrators. A closer look at data of our previous fMRI study on parochialism (Baumgartner et al., 2012) might explain the observed direction of the TMS finding. With a less stringent significance threshold, we found correlative evidence (at \( P < 0.05 \), \( r = 0.498 \), \( \text{ETA}^2 = 25\% \)) that the functional connectivity between right TPJ and DMPFC is linked to the punishment of outgroup perpetrators. The stronger the functional connectivity between these two areas, the more severely outgroup perpetrators are punished (see Supplementary Figure 6). This might explain the direction of the observed TMS effect, namely that TMS over the right TPJ reduced the punishment of outgroup perpetrators.

Certain regions nearby the TPJ (e.g. in the inferior parietal cortex) are known to be involved in attention and processing of more general task performance (Corbetta et al., 2000; Mitchell, 2008; Scholz et al., 2009; Van Overwalle, 2011). Thus, the question arises whether our findings are rather unspecific and due to diminished attention or generally diminished task performance. We believe that this is unlikely for the following reasons: First, the rTMS effect is only observable in the condition where perpetrators defected against cooperating interaction partners (DC), but not in conditions where there was no reason to punish (i.e. in the case of cooperative behavior in the PDG). Moreover, we only find a rTMS effect in the intergroup constellation, i.e. when the social category of the perpetrator to whom punishment could be assigned is different, but not in the intragroup constellation, i.e. where both perpetrators belong to the same social category, namely the ingroup. Such highly specific findings would be difficult to explain with unspecific attentional or cognitive processes because they would be more likely to affect all punishment patterns and conditions in a similar way. Second, we find no treatment group effects on response times. Repeated-measures ANOVA with within-subject factors behavioral trials (CC, CD, DC and DD), group constellations (OUT/IN, IN/OUT, IN/IN) and between-subject factor treatment groups (left TMS, right TMS, sham) revealed neither a main effect of treatment groups nor interaction effects of treatment groups with one or both other factors (all \( P > 0.487 \)). These findings are not consistent with the assumption of an unspecific effect of attention or more general task performance because such an effect would rather manifest itself in slower response times.

Summing up, this is the first study providing causal evidence that the right TPJ plays a pivotal role in modulating parochialism in intergroup conflict. Based on the evidence of decreased parochialism after right TPJ disruption together with the observation that the differential behaviors toward members of different social groups often underlie acts such as discrimination, it is tempting to conclude that a disrupted TPJ could lead to the world being a better place. Such diminished parochialism is only evident, however, if the brain is not functioning properly. Therefore—although parochialism in its extreme can lead to severe and detrimental outcomes (Haushofer et al., 2010)—treating an ingroup member differently than an outgroup member appears to have a necessary function that has developed through evolution. As a social species, individuals have evolved living in different social groups and cognitive adaptations must therefore have developed to differentially treat individuals according to group membership (Efferson et al., 2008). It thus seems plausible that along with the cognitive adaptations, evolution has engineered neural adaptations to distinguish the behavior of ingroup and outgroup members and to differentiate the punishment of norm transgressions depending on the perpetrator’s social category. Our study suggests that these neural adaptations have developed in the right TPJ.

SUPPLEMENTARY DATA

Supplementary data are available at SCAN online.

Conflict of Interest

None declared.

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