Garlic intake and the risk of colorectal cancer
A meta-analysis

Xi Zhou, PhD, Hailua Qian, PhD, Dan Zhang, PhD, Li Zeng, PhD

Abstract
This was a meta-analysis of epidemiological articles that aimed to estimate the association of garlic intake with the risk of colorectal cancer (CRC).

Electronic databases, including the Cochrane Database of Systematic Reviews, PubMed, and EMBASE, were systematically searched from inception to May 2019 to identify related articles. In addition, a random model was used to pool the included evidence based on heterogeneity. Additionally, subgroup analyses were carried out to examine the differences between different groups. The stability of our findings was tested through sensitivity analyses. Publication bias was also assessed by Egger and Begg tests. Moreover, all enrolled studies were ordered according to the publication year for a cumulative meta-analysis.

A total of 11 studies (involving 12,558 cases) were included in the current meta-analysis. Our integrated relative risk (RR) of CRC was 0.80 [95% confidence interval (CI), 0.69–0.91] for the highest versus the lowest garlic consumption categories (RR: 0.71 [95% CI, 0.60–0.84] for controls and RR: 0.99 [95% CI, 0.80–1.23] for cohorts). There was significant heterogeneity across all enrolled studies (I² = 68.3%, P < .01). The sensitivity analysis revealed no notable alterations of the integrated results. According to the funnel plot regarding garlic intake and the risk of CRC, together with the Egger test (P = .1) and Begg test (P = .064) results, there was no notable evidence of publication bias. The cumulative meta-analysis suggested that the 95% CIs became narrower with the increase in sample size.

Based on the existing evidence, garlic intake could reduce the risk of CRC.

Abbreviations: CI = confidence interval, CRC = colorectal cancer, NA = not available, NR = not reported, RR = relative risk.

Keywords: colorectal cancer, diet, epidemiology, garlic, meta-analysis

1. Introduction
Colorectal cancer (CRC) has the third highest rate of morbidity worldwide, behind lung cancer and breast cancer. Altogether, 1,849,518 patients are diagnosed with CRC every year, accounting for 10.2% of all cancer cases, and CRC ranked second and third in prevalence among females and males, respectively, in 2018.[1] In addition, CRC is the second leading cause of cancer-related death worldwide, regardless of certain geographical differences with regard to its morbidity and mortality. Typically, Asia has the greatest morbidity (n = 957,896; 51.8%) and mortality (n = 461,422; 52.4%), irrespective of age and sex. Many scientific articles have suggested that the consumption of vegetables may impact the incidence of cancer; in particular, certain geographical and environmental factors, including diet, may play a vital role in CRC development.[2–4]

Garlic has been used globally in numerous ways. According to the World Cancer Research Fund/American Institute for Cancer Research, garlic consumption has a potentially favorable effect against CRC.[5] Some evidence has suggested that garlic consumption can lower CRC risk.[6–8] Related data from the past 10 years have updated the knowledge in this regard, and some epidemiological studies have examined the relationship of garlic intake with CRC risk; nonetheless, no definite conclusion can be made. Two meta-analyses published in 2014[9,10] quantitatively estimated the association and suggested that garlic showed no evident effect in terms of CRC risk reduction. In contrast, Federica Turati’s article[6] based on 9 articles published in the same year showed a pooled RR of CRC of 0.85 [95% confidence interval (CI), 0.72–1.00), and it was concluded that garlic intake might protect against CRC. In addition, a hospital-based case-control study carried out in Shanghai found that garlic consumption was linked to a reduced risk of CRC.[11] Additionally, the American Cancer Society Cancer Prevention Study II cohort indicated that CRC risk was weakly inversely associated with dietary garlic intake among females, but such a risk was possibly increased among males.[12] Additionally, an Italian and Swiss case-control study found that garlic intake could protect against CRC.[13] To further explore the relationship between garlic intake and CRC risk, the current meta-analysis was carried out.

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2. Materials and methods

2.1. Retrieval strategy and screening criteria

Articles examining the relationship of garlic intake with CRC were systematically retrieved from the Cochrane Database of Systematic Reviews, PubMed and EMBASE databases from inception through May 2019 by 2 researchers (Xi Zhou and Dan Zhang) independently. The following string was used for the retrieval: “(((garlic OR allium sativum) OR vegetable) AND (((cancer) OR neoplastic disease) OR neoplasm) OR tumor) OR carcinoma) AND (((rectal) OR colon) OR colorectal).” Moreover, these 2 researchers were also responsible for retrieving and assessing the candidate studies that examined the relationship of garlic intake (excluding supplementary intake of garlic) with the risk of CRC based on human observational research published in the English language. Additionally, the reference lists of the selected studies were also screened to avoid omitting any eligible study. If we required more information, the original authors were contacted.

The study inclusion criteria were as follows:
(1) cohort or case-control studies;
(2) studies evaluating the relationship of garlic consumption with CRC risk;
(3) studies reporting the relative risk (RR) or odds ratio (OR) and 95% CI or the essential data for calculation.

For repeated studies carried out among identical research populations, the most informative study was included. Two studies were excluded because they examined garlic supplementation rather than dietary garlic intake. In addition, studies examining pepper and onion consumption together with garlic intake or the overall intake of allium vegetables were excluded. In addition, 2 studies were carried out with an identical study population, and the most informative one was included.

2.2. Data collection and quality evaluation

Three researchers (Xi Zhou, Dan Zhang, Haihua Qian) were responsible for collecting and verifying the data. Any disagreement among them was settled by mutual negotiation. Data were collected from all included studies, including the last name of the first author, year of publication, study design, study region/country, sex, subject number (including cases and controls/noncases/cohort size), and estimated RR (for >1 estimated RR in 1 study, the I^2 adjusting for the most confounding factors was included in the analysis) with the 95% CI (the greatest versus smallest garlic consumption categories; garlic supplement users included). When necessary, the data of cases and noncases/persons that were at risk were adjusted for every exposure category and covariate before subsequent analysis.

Study quality was evaluated based on the 9-star Newcastle–Ottawa scale (NOS) independently by 2 researchers (Xi Zhou and Dan Zhang). Each study was assigned 9 stars at most, and studies assigned ≥7 stars were of high quality.

2.3. Statistical analyses

In this study, RRs were utilized as a risk estimate. There is a low absolute CRC risk among humans; as a result, ORs were directly deemed as RRs. The fixed-effects model was utilized when studies reported the colon and rectal cancer risks separately to integrate all risk estimates while obtaining the overall CRC estimate. The I^2 statistic and Cochrane Q test were adopted to assess the potential heterogeneity among the included studies. Moreover, the fixed-effects model was employed in the absence of distinct heterogeneity; otherwise, the random-effects model was adopted. Sensitivity analysis was also carried out by removing each individual study 1 at a time in sequence to assess the effect of each study on integrated risk estimates. Subgroup analyses were also carried out as stratified based on the cancer position, sex, geographic area, sample size, and study design. Begg and Egger linear regression was employed to evaluate publication bias. Differences with P < .05 were regarded as statistically significant. STATA (version 12; StataCorp, College Station, TX) was utilized for all statistical analyses.

2.4. Ethical statement

This is a meta-analysis, which was based on previous published studies and did not have original data. Therefore, no ethical approval and patient consent are required.

3. Results

3.1. Literature retrieval

Altogether, 5081 candidate studies were identified during the initial retrieval, among which 5049 were eliminated after abstract or title screening. For the remaining 32 studies, the full texts were carefully read. One candidate study were identified after viewing the reference lists of the included studies. After full-text assessment, 11 articles were retrieved and included in our meta-analysis. Figure 1 presents the study retrieval process.

3.2. Study characteristics

A total of 11 articles published from 1991 to 2018 involving 12 datasets (n=12,558) were included in the current meta-analysis. Among these studies, the article from Meng et al recruited 2 cohorts, namely, the Health Professionals Follow-up Study (HPFS) and the Nurses’ Health Study (NHS). Table 1 displays the major features of the 11 articles. We defined studies with a sample size ≤10,000 as having a small sample size; otherwise, they had a large sample size. Eight of the 11 studies had small sample sizes (n=35,216–99,700), Furthermore, these studies were conducted in many regions, among which 6 datasets were from America, 2 datasets were from Asia, 3 datasets were from Europe, and 1 dataset was from Australia. The number of CRC cases among these articles ranged from 109 to 2280. Table 1 shows the study quality score on the basis of the 9-star NOS system. Two studies had 7 stars, 6 studies had 8 stars, and 4 had 9 stars. Based on the quality evaluation criteria, each of the studies enrolled in the current meta-analysis were of high quality. Also, we conducted Grading of Recommendations Assessment Development and Evaluation (GRADE) evaluation of the evidence in Table 2.

3.3. Overall and subgroup analysis

Eleven studies mentioned risk estimates regarding garlic consumption and CRC risk. The random-effects model was used in the meta-analysis, which revealed an increased CRC risk
Table 1
Main characters of included studies.

| Reference                  | Study                                                                 | Country          | Number of participants | Male | Female | Total | Study quality | Design     |
|----------------------------|----------------------------------------------------------------------|------------------|------------------------|------|--------|-------|---------------|------------|
| Hu et al (1991)            | The First Hospital of Harbin City and the General Hospital of Forestry | China            | NA                     | 218 (109 cases) | 218 (109 cases) | 7     | Case-control  |            |
| Witte et al (1996)         | The Southern California United States                                | United States    | 488 (325 cases)        | 488 (163 cases) | 976 (488 cases) | 8     | Case-control  |            |
| Le Marchand et al (1997)   | The National Cancer Institute Surveillance, Epidemiology, End Results (SEER) Program | United States    | 1192 (698 cases)       | 1192 (494 cases) | 2384 (1192 cases) | 9     | Case-control  |            |
| Silvia et al (1997)        | The 6 Italian areas                                                  | Italy            | 5155 (1125 cases)      | 5155 (928 cases) | 5155 (1953 cases) | 8     | Case-control  |            |
| Levi et al (1999)          | University Hospital of Lausanne Northern Italy                      | Switzerland      | 714 (142 cases)        | 714 (81 cases)  | 714 (223 cases)  | 9     | Case-control  |            |
| Galeone et al (2006)       | The Western Australian Bowel Health Study                            | Australia        | NR                     | NR               | 4017 (1770 cases) | 8     | Case-control  |            |
| Annema et al (2011)        | The First Hospital of China Medical University                      | China            | 1666 (433 cases)       | 1666 (350 cases) | 1666 (833 cases) | 9     | Case-control  |            |
| Wu et al (2018)            | The Iowa Women’s Health Study                                        | United States    | NA                     | 35,216 (212 cases) | 35,216 (212 cases) | 7     | Cohort        |            |
| Steinmetz et al (1994)     | The American Cancer Society Cancer Prevention Study (CPS) II Nutrition Cohort | United States    | 42,824 (579 cases)     | 56,876 (551 cases) | 99,700 (1130 cases) | 9     | Cohort        |            |
| Meng et al (2013)          | The Nurses’ Health Study (NHS)                                       | United States    | NA                     | 76,208 (1339 cases) | 76,208 (1339 cases) | 8     | Cohort        |            |
| Meng et al (2013)          | The Health Professionals Follow-up Study (HPFS)                      | United States    | 45,592 (1029 cases)    | NA               | 45,592 (1029 cases) | 8     | Cohort        |            |

NA=not available, NR=no reported.
Garlic consumption was related to a lower risk of CRC (Fig. 2). As shown in Table 3, subgroup analyses were carried out to examine the underlying factors associated with CRC risk between the highest versus lowest garlic consumption groups. As observed, garlic consumption had a stronger association with rectal cancer than with colon cancer. European studies suggested a stronger inverse relationship of garlic than that identified in Asian and American studies, which was especially notable in the only 1 Australian study. Additionally, the sample size of the studies might have affected our assessment of the relationship of garlic intake with CRC. The small sample size suggested that garlic intake reduced CRC risk. Based on the study design, there was relevant heterogeneity among articles, and no definite results

Table 2

| GRADE evidence profile of garlic intake for the risk of colorectal cancer. |
|---------------------------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|
| Certainty assessment           | No of patients  | Effect          |                  | Other considerations |
|---------------------------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|
| No of studies                   | Study design    | Risk of bias    | Inconsistency   | Indirectness    | Imprecision     | With no garlic | With garlic intake | Relative (95% CI) | Absolute (95% CI) | Certainty | Importance |
| 8                               | Observational studies (Case-control) | Not serious | Not serious | Not serious | None | 8846 cases | OR 0.71 (0.60–0.84) |              |              | LOW      | CRITICAL   |
|                                 |                 |                 |                 |                 |                 | 13,327 controls |               |              |              |            |            |
| 4                               | Observational studies (Cohort) | Not serious | Not serious | Not serious | None | 3710/256,716 | OR 0.99 (0.80 to 1.23) |              |              | LOW      | CRITICAL   |
|                                 |                 |                 |                 |                 |                 | 253,006/256,716 |               |              |              |            |            |

CI = confidence interval, OR = odds ratio.

(RR: 0.80; 95% CI: 0.69–0.91). Garlic consumption was related to a lower risk of CRC (Fig. 2).

As shown in Table 3, subgroup analyses were carried out to examine the underlying factors associated with CRC risk between the highest versus lowest garlic consumption groups. As observed, garlic consumption had a stronger association with rectal cancer than with colon cancer. European studies suggested

![Forest plots for the associations between garlic intake and colorectal cancer risk.](image)

Figure 2. Forest plots for the associations between garlic intake and colorectal cancer risk.
were obtained from cohort studies (Fig. 3). In subgroup analyses stratified by sex, no statistical significance was detected in the male and female groups.

### 3.4. Heterogeneity assessment

There was significant heterogeneity among the included articles ($I^2 = 68.3\%$, $P < .01$). As a result, the Galbraith plot test was conducted to explore the possible heterogeneity source based on the included articles. Nonetheless, our results suggested that 3 articles [11,28,33] might have been the leading sources of heterogeneity. No significant change was detected when these 3 studies were excluded (RR: 0.80; 95% CI: 0.71–0.91).

### 3.5. Cumulative meta-analysis

All enrolled articles were ordered according to their year of publication to conduct a cumulative meta-analysis. According to the cumulative meta-analysis results, the relationship of garlic consumption with CRC risk was in chronological order (Fig. 4). In addition, the corresponding 95% CIs became narrower as the

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**Table 3**

Summary relative risks of colorectal cancer and corresponding 95% confidence intervals for the highest versus the lowest category of garlic intake of selected factors.

| Subtype          | n studies | RR (95% CI)       | P heterogeneity; $I^2$ |
|------------------|-----------|-------------------|------------------------|
| Colon            | 6         | 0.87 (0.72, 1.06) | .009 67.3%             |
| Rectum           | 6         | 0.75 (0.57, 1.00) | .035 58.2%             |
| Gender           |           |                   |                        |
| Female           | 6         | 0.75 (0.54, 1.04) | .009 67.7%             |
| Male             | 4         | 0.86 (0.65, 1.13) | .104 51.3%             |
| Geographic area  |           |                   |                        |
| America          | 6         | 0.88 (0.71, 1.09) | .023 61.6%             |
| Asia             | 2         | 0.44 (0.19, 1.01) | .196 42.8%             |
| Europe           | 3         | 0.74 (0.58, 0.96) | .003 82.6%             |
| Australia        | 1         | 0.86 (0.68, 1.09) | – –                    |
| Sample size      |           |                   |                        |
| <10,000          | 8         | 0.71 (0.60, 0.84) | .002 69.7%             |
| >10,000          | 4         | 0.99 (0.80, 1.23) | .127 47.2%             |
| Study design     |           |                   |                        |
| Case-control     | 8         | 0.71 (0.60, 0.84) | .002 69.7%             |
| Cohort           | 4         | 0.99 (0.80, 1.23) | .127 47.4%             |

CI = confidence interval, RR = relative risk.

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**Figure 3.** Subgroup analysis of study design for the associations between garlic intake and colorectal cancer risk.

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CI = confidence interval, RR = relative risk.
sample size increased, suggesting increasing estimated accuracy with the continuous inclusion of studies.

3.6. Sensitivity analysis

The influence of 1 article on the overall estimates was assessed through sensitivity analysis, which was performed by eliminating 1 individual article at a time in sequence. According to the results of the sensitivity analysis, the RR varied from 0.76 (95% CI: 0.67–0.88), when the study from Meng et al[28] was excluded [NHS] to 0.82 (95% CI: 0.72–0.93) when the study from Levi et al[33] was eliminated (Fig. 5). There was no notable difference among the pooled findings.

3.7. Publication bias

The potential publication bias among the enrolled articles was assessed by Egger and Begg tests (Fig. 6). There was no notable evidence of publication bias (P = .1 and P = .064, respectively).

4. Discussion

The current meta-analysis integrated the findings from 11 epidemiological studies, which included 8 case-control and 4 cohort studies (Meng [NHS] and Meng [HPFS] were regarded as 2 cohort studies). Our results suggested that garlic consumption was associated with protection against CRC (RR = 0.80, 95% CI: 0.69–0.91).

The extent of heterogeneity among all included articles was assessed through a quantified I² test and Q test, and significant heterogeneity was detected (I² = 68.3%, P < .01). A Galbraith plot test was carried out, which revealed that 3 studies might be the main sources of heterogeneity, but no notable change was found when excluding these studies (RR = 0.80, 95% CI: 0.71–0.91). However, there was heterogeneity among the included articles, especially based on various regions and study design. The results of the case-control (RR = 0.71, 95% CI: 0.60–0.84) and cohort (RR = 0.99, 95% CI: 0.80–1.23) studies are shown. The following reasons may explain these results. First, this meta-analysis included observational studies, which might be influenced by different sources of bias, especially in regard to the retrospective studies. For a retrospective study, recall bias may partially lead to case-control results that are different from cohort results, which is ascribed to the influence of recent CRC diagnosis on recalling dietary habits. In addition, for case-control studies, health-conscious cases may have overreported the consumption of “health food” (such as vegetables), which may thereby hinder
the potential determination of an actual relationship. Second, all cohort studies were carried out in America, a place distinct from Europe. In Europe, local people generally eat olive oil, tomatoes, and garlic in sauces of pasta and salads.[35] The per capita garlic consumption in the US has increased greatly over the past few decades.[6] In conclusion, this meta-analysis suggested that greater garlic intake led to a reduced CRC risk.

The cancer-prevention mechanism of garlic remains unclear. S-allylmercaptocysteine, the water-soluble derivative of garlic, has been found to display anti-proliferative capacity in numerous cancer cell lines, and it has also shown tumor-inhibiting effects in in vivo conditions.[36,37] Garlic can enhance the anticancer ability by activating the antioxidant transcription expression of Nrf2 and the downstream gene NQO1.[38] Aged black garlic (ABG) can modify the mRNA expression of neuropeptides and proteins in the case of inflammation.[39] Alternatively, ABG can exert its anticancer function by suppressing the proliferation of cells.[40] Garlic is potentially potent against some types of cancers; importantly, it is a universal anticancer drug that is potent against different cancers[41]; this effect is achieved by suppressing the
growth of cancer cells and effectively suppressing the proliferation of infiltrative macrophages within the tumor-like microenvir-

onment. A large amount of preclinical data suggests that garlic has a certain effect on modulating the metabolism of carcinogens, inhibiting the progression of the cell cycle, inducing apoptosis and histone modification, and inhibiting angiogene-
sis. Studies in vivo and in vitro suggest that garlic, together with garlic compounds, can protect against carcinogenesis by forming the carcinogen-DNA adduct, through angiogenesis and cell proliferation, by promoting apoptosis and anti-tumorigenic genes and redistributing the cell cycle.

Some limitations should be noted in the current meta-analysis. First, all data were derived from observational studies, which resulted in residual or unmeasured confounding factors, even though the identified factors were adjusted in many included articles. Second, the unit and definition of garlic intake in every category were not standardized across all included studies. Third, certain subgroup analyses had a small dataset. Fourth, with regard to the features of case-control and cohort studies, the latter is a superior approach for illustrating the relationship of garlic consumption with CRC risk. In addition, the results obtained based on cohort studies are reliable and valuable because they are controlled for confounders. Finally, only the English articles were included in the current meta-analysis, which might lead to possible publication bias, even though Egger test revealed no evident publication bias.

Nonetheless, this meta-analysis had certain strengths. Specifically, the reporting recommendations from the MOOSE group were followed, the retrieval string was used, and concept terms including “vegetable” and “allium sativum” were used, rather than depending on specific terms associated with garlic. This made it possible to identify studies that provided data regarding garlic from abstracts or titles, as well as those providing data from tables or the main text.

5. Conclusions

In summary, garlic intake can lower CRC risk. Nonetheless, such a conclusion was mainly drawn based on case-control studies, while no definite result was obtained from cohort studies. Thus, our results should be interpreted with caution. More epidemiological results, especially large-scale prospective cohort studies, can contribute to elucidating and quantifying the underlying protective effect of garlic against CRC.

Author contributions

Data curation: Xi Zhou, Haihua Qian, Dan Zhang.
Methodology: Xi Zhou, Li Zeng.
Software: Xi Zhou, Dan Zhang.
Supervision: Haihua Qian, Dan Zhang.
Writing – original draft: Xi Zhou.
Writing – review and editing: Li Zeng.

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