The relationship between disability and parental status: A register study of the 1968 to 1970 birth cohorts

Fredinah Namatovu (fredinah.namatovu@umu.se)
Umea Universitet https://orcid.org/0000-0001-5471-9043

Erling Häggström Lundevaller
Umea Universitet

Lotta Vikström
Umea Universitet

Research article

Keywords: Child, children, disability, disability benefits, disability pension, early retirement, fertility, having children, parenthood, parenting, parental status

DOI: https://doi.org/10.21203/rs.3.rs-61041/v3

License: © This work is licensed under a Creative Commons Attribution 4.0 International License. Read Full License
Abstract

Background Having children is a major life course event yet some disabilities could make it biologically challenging and some others could limit access to necessary socioeconomic resources. To date, there is relatively little data on disability and parental status and our study aims to investigate this relationship.

Methods This longitudinal cohort study was based on register data obtained from all people born in Sweden from 1968 to 1970 (n=440,202). We performed descriptive analyses, graphical plots, logistic regression and Cox regression analyses.

Results Our findings from both logistic regression and Cox regression indicated that individuals that started to receive disability benefits at an early age had reduced chances of having children during the follow-up duration. Men with disabilities were less likely to have children when compared to women with disabilities and to men and women without disabilities.

Conclusions We found evidence that disability during early adulthood was associated with reduced chances of having children. Findings support policies and programmes aimed at promoting optimal health during early adulthood, as this would promote continued labour force participation, reduce early use of disability benefits and possibly improve chances of becoming a parent.

Background

Many societies consider having children a major life course event [1]. In modern societies, children are shown to offer psychosocial benefits, such as the intrinsic pleasure derived from watching one's own children grow and having someone for regular interaction [2]. Having children is also associated with improved physical health and enhanced health behaviours, such as cessation of smoking and alcohol abuse [3]. However, the decision to have or not to have children is based on an individual's current and future socio-economic circumstances [1, 4]. Several researchers have shown that fertility patterns are fundamentally affected by social, cultural and economic factors [5, 6]. In high income economies, there are a number of pivotal factors that are often fulfilled before starting to have children. Such factors include being out of school, having a steady income, living independently and having a partner [5, 6].

Having a disability can profoundly affect one's ability to fulfil the pivotal determinants for having a first child and might lead to postponement or in some cases to childlessness. Some disabilities make it biologically impossible to have children [7], while other disabilities might limit access to the required socioeconomic resources [8]. Literature operationalises disabilities differently, some definitions include self-reported disability, disability diagnosis and the administrative definition. In this study, we use an administrative definition that considers one to have a disability if a person receives disability benefits [9-11]. In Sweden, disability benefits are part of a public social security programme that provides income support to people of working age that experience long-term health limitations in their working capacity [9-11]. Having a disability that requires exiting the work force at a young age increases the length of time spent outside of a work environment, which can jeopardise one's health, social and economic conditions over their lifetime [12]. European data suggests that the number of young people exiting the labour market early is increasing [12]. However, there is limited research on whether disability that leads to early exit from the labour force is associated with reduced chances of having children.
From a theoretical point of view, the link between disability and parental status can be understood using Oppenheimer’s uncertainty hypothesis [13, 14] and the theory of assortative mating [15, 16]. Oppenheimer’s uncertainty hypothesis suggests that insufficient economic resources hinder marriage [13, 14]. Exiting the labour force and being on disability benefits early in life might create financial constraints and economic uncertainty. Oppenheimer also argues that assortative mating might occur when individuals with low economic resources are considered undesirable for marriage [13, 14]. Due to assortative mating on disability, people without disabilities might prefer to choose partners that do not have disabilities. Disability researchers have used the theory of assortative mating to explain why disability is linked to reduced chances of marriage and cohabitation [15-18]. Disability that occurs during early adulthood could contribute to assortative mating, creates economic constraints, makes it difficult for one to have a steady income, live independently and have a partner, which results in the postponement of parenthood or even childlessness.

Current literature suggests that people with disabilities report limited economic opportunities and high poverty rates [8, 19]. Several studies also indicate that disability is associated with low levels of partnership and high levels of single living [17, 18, 20-22]. People with disabilities also report negative societal attitudes characterised by infantilisation and being treated as asexual [22-24]. Structural barriers and limited access to fertility services are additional challenges reported in this population [25, 26]. All these factors negatively impact the chances of having children among people with disabilities.

Compared to the extensive body of research on parenthood in the general population, there is relatively little data on the parental status of people with disabilities. Available research can be divided into two subsets. One body of research examines the experiences of parenting with a disability and being a parent to a child with a disability [22-28]. Another branch of research focuses on the parenthood status of people with learning disabilities and physical disabilities [25-28]. There is still little research on the extent to which people with disabilities become parents and the association between disability and parental status.

Aims
The aim of this study was to examine the relationship between having a disability in early adulthood and reduced chances of having children.

Methods
Design and study setting
In this longitudinal cohort study, we used national register data that consisted of all people that were born in Sweden (n=440,202) from 1968 to 1970. Individuals were followed from birth until 2010 when the youngest were 39 years old and the oldest were 42 years old. The selection of this study duration was based on data availability. All data was obtained from the Longitudinal Integration Database for Health Insurance and Labour Market Studies (LISA database). Our data consisted of information on total population, disability benefits status, number of children, sex, and year of birth. All data was anonymised by Statistics Sweden and was made available for analysis through the Swedish Initiative for Research on Microdata in Social and Medical Sciences (Umeå SIMSAM Lab) [29].
Outcome variable

The outcome variable was whether an individual became a parent or not from age 23 up to end of the study duration. The variable “becoming a parent” was considered binary, which indicated whether an individual had a biological child or not during the observation time.

Exposure variables

The main exposure variable to indicate disability status was “whether an individual received a disability benefit at age 20–22 years or not.” Eligibility for receipt of these benefits was confirmed through a medical examination which indicated diminished health and work capacity [9-11]. During the study duration, the disability benefits were referred to as a disability pension and were awarded to medically eligible people aged between 16 and 64 years [9-11]. We coded the individuals who received disability benefits “yes” at the first year they received the benefits; those who did not receive disability benefits during the entire study duration were coded “no”. For this study, age at first receipt of disability benefits ranged between 20 and 22 years. Those aged 20 years could have started on the benefits earlier than this, however this was the age when they were first registered in the LISA database following its establishment in 1990. The other covariate included was the year of birth, which ranged between 1968 and 1970.

Statistical analyses

Descriptive statistics were performed using frequency tables to give an overview of the dataset. We plotted the Kaplan-Meier curve (Figure 1) to give a depiction of the time to the estimated probability of having a child.

Logistic regression analyses were performed for men and women separately to assess the association between having a disability at ages 20–22 and later having a child. Independent models for disability benefits and year of birth and models that simultaneously included the two variables were reported (Table 2). Regression results were presented as odds ratios (ORs) with 95% confidence intervals (CIs) and statistical significance set at p < 0.05.

Using Cox proportional hazard regression, we modelled having a child during the follow-up. Disability status was assessed at age 20–22 years, and age at having a first child set at 23 years and onwards to ensure the temporal order in this longitudinal data. Individuals entered the study at age 20–22 and were followed until they had their first child or until they were censored, out migrated, died, or the end of study period on 31, December 2010. We reported four models, which included independent and adjusted associations fitted for men and women separately. In the results presented in the Appendix, we repeated both the independent and adjusted logistic regression and Cox regression analyses incorporating marital status, to disentangle the effect of marriage. We further tested the interactions between disability and marital status. All analysis was performed using R software 4.0.2 for windows.

Results
For the Swedish birth cohorts of 1968 to 1970, childlessness was more than double among people with disabilities compared to those without disabilities. For those with a disability, 13% had at least one child compared to 73% without a disability. The number of men was slightly higher than that of women (51% vs. 49%). The 1968 birth cohort had the highest proportion with disabilities (Table 1).

Table 1. The distribution of Swedish men and women born between 1968–1970 by selected demographic characteristics presented for those with or without disability at age 20–22 years (N=440,221)

| Descriptive          | No Disability (n=437,367) | Disability   | Total (440,220) |
|----------------------|---------------------------|--------------|-----------------|
|                      | N (%)                     | N (%)        | N (%)           |
| Any children         |                           |              |                 |
| No                   | 119,530 (27)              | 2,488 (87)   | 122,018 (28)    |
| Yes                  | 317,837 (73)              | 365 (13)     | 318,202 (72)    |
| Number of children   |                           |              |                 |
| 1                    | 64,410 (15)               | 139 (05)     | 64,549 (15)     |
| 2+                   | 253,427 (58)              | 226 (08)     | 253,653 (58)    |
| Age at first child   |                           |              |                 |
| No Child             | 119,530 (27)              | 2,488 (87)   | 122,018 (28)    |
| 13–20                | 20,252 (05)               | 36 (01)      | 20,288 (05)     |
| 21–25                | 88,425 (20)               | 123 (04)     | 88,548 (20)     |
| 26–30                | 96,833 (22)               | 98 (03)      | 96,931 (22)     |
| 31–35                | 76,196 (17)               | 65 (02)      | 76,261 (17)     |
| 36–42                | 36,131 (08)               | 43 (02)      | 36,174 (09)     |
| Sex                  |                           |              |                 |
| Men                  | 222,989 (51)              | 1,592 (56)   | 224,581 (51)    |
| Women                | 214,260 (49)              | 1,261 (44)   | 215,521 (49)    |
| Year of index birth  |                           |              |                 |
| 1968                 | 149,709 (34)              | 829 (29)     | 150,538 (34)    |
| 1969                 | 142,943 (33)              | 927 (32)     | 143,870 (33)    |
| 1970                 | 144,715 (33)              | 1,097 (38)   | 145,812 (33)    |

*had a child by end of follow-up, #Disability present at age 20-22, ¢ever married during the study duration

Figure 1 shows that childlessness was more common among those with disabilities. Childlessness rapidly decreased with time in the group without disabilities but remained high and barely changed in the group with disabilities. At age 30, about 80% of the women and 95% of men with disabilities were childless, with meagre fluctuations by the end of the follow-up. Corresponding figures among those without disabilities were 50% for women and about 65% for men at age 30, this further dropped to about 10% in women and over 25% in men.
Regression results

The bivariate and multivariable logistic regression results presented in Table 2, showed that men and women that received their first disability benefits before their 23rd birthday were significantly less likely to become parents compared to those that never received disability benefits. In addition, men with disabilities were less likely to become parents compared to women with disabilities. The birth cohorts of 1969 and 1970 had lower chances of having children, compared to the birth cohort of 1968, regardless of sex.

Table 2. Logistic regression results on the association between receiving disability benefits at 20–22 years and having a first child at a later age.

| Variable                      | Men |       |       | Women |       |       |
|-------------------------------|-----|-------|-------|-------|-------|-------|
|                               | COR (95% CI) | AOR (95% CI) | COR (95% CI) | AOR (95% CI) |       |       |
| Disability benefits: No       | 1.0 | 1.0   | 1.0   | 1.0   |       |       |
| <23yrs                        | 0.02 (0.02–0.03) | 0.02 (0.02–0.03) | 0.04 (0.03–0.04) | 0.04 (0.03–0.04) |       |       |
| Year of birth                 |     |       |       |       |       |       |
| 1968                          | 1.0 |       |       |       | 1.0   |       |
| 1969                          | 0.97 (0.94–1.00) |       | 0.96 (0.92–0.99) |       |       |
| 1970                          | 0.93 (0.90–0.96) |       | 0.95 (0.92–0.99) |       |       |

COR=Crude Odds Ratio; AOR=Adjusted Odds Ratio; CI=Confidence Interval

The hazard ratio from the Cox regression presented in Table 3 also confirmed results from logistic regression, disability was significantly associated with reduced chances of having children, regardless of sex. Men with disabilities were less likely to become parents compared to women with disabilities. Again, becoming a parent was less common in younger birth cohorts compared to the birth cohort of 1968.

Table 3. The hazard ratio of having a first child if you received disability benefits at 20–22 years.

| Variable                     | Men |       |       | Women |       |       |
|------------------------------|-----|-------|-------|-------|-------|-------|
|                               | CHR (95% CI) | AHR (95% CI) | CHR (95% CI) | AHR (95% CI) |       |       |
| Disability benefits: No      |     |       |       |       | 1.00   | 1.00   |
| at 20–22yrs                  | 0.06 (0.05–0.07) | 0.06 (0.05–0.07) | 0.11 (0.09–0.13) | 0.11 (0.09–0.13) |       |       |
| Year of birth                |     |       |       |       |       |       |
| 1968                         | 1.0 |       | 1.00  | 1.00  | 1.00  |       |
| 1969                         |       | 0.98 (0.97–0.99) |       | 0.96 (0.95–0.98) |       |       |
| 1970                         |       | 0.95 (0.94–0.97) |       | 0.95 (0.93–0.96) |       |       |

CHR=Crude Hazard Ratio; AOR=Adjusted Hazard Ratio; CI=Confidence Interval
The results presented in the Appendix, also confirmed that disability was significantly associated with reduced chances of having children even after adjusting for marital status and assessing the interaction between disability and marriage. This association was confirmed in both the logistic regression and Cox regression models, Tables 4, 5, 6 and 7.

**Discussion**

Our register data for the birth cohorts of 1968 to 1970 showed that having a disability during early adulthood was associated with extremely low chances of having children during the observed duration. Men with disabilities were less likely to have children compared to women with disabilities and to men and women without disabilities. This study also noted that the chances of having children were reduced for younger birth cohorts compared to the birth cohort of 1968.

Our finding that receiving disability benefits was associated with significant reduction in chances of having children support findings from previous studies that reported similar observations [20, 21]. The association between disability and low chances of having children could very well be explained by several factors including economic barriers as Oppenheimer’s theory suggests [13, 30] and by assortative mating on disability status [15, 16]. In addition, biological barriers and negative societal attitudes on parenting with a disability could also negatively influence the likelihood of becoming a parent [24, 25].

Having a disability in early adulthood was associated with a significant reduction in the chances of becoming a parent later in life. Starting to receive disability benefits during the early twenties might imply prolonged ill health and socio-economic constraints, factors that might constrain the decision to become a parent. Moreover, previous data has shown that the majority of individuals who start to receive disability benefits continue to do so on a long-term basis, with an outflow rate at just about three percent [12].

This study also identified major sex differences in the chances of having children. The chances were much lower for men with disabilities compared to women with disabilities or to men and women without disabilities. Literature suggests that entry into a stable partnership requires a strong financial underpinning and more occupational stability for men compared to women [13, 14]. Limited male entry into partnership formation will subsequently result in low numbers of children for men. However, it is also important to note that our study population consisted of a higher proportion of men than women, implying a surplus of men on the partnership market, which further constrain the partnership chances of men with disabilities.

We noted that the odds of having children were reduced in the younger birth cohorts of 1969 and 1970 compared to people born in 1968. There are no known social economic changes during this time to explain these parental status differences by birth cohort. It is therefore possible that this association is due to random variation.

**Strengths and limitations of the study**
The main strength of this study is the access to a large longitudinal dataset that gave us an opportunity to observe a total population for a long duration. Moreover, using register data ensured minimal loss to follow-up, data on the use of disability benefits and having children was available for all cohorts. In this study, we define disability status based on use of disability benefits, we consider this a reliable measure for disability status and the severity of the disability given that eligibility was based on medical evidence suggesting long-term health limitations in working capacity. Our findings should be interpreted within the frames of the studied type of disability. Further research using other definitions of disabilities could help clarify the observed association between disability and chances of having children. Even though our models excluded some people due to missing data, there is no reason to suspect that missing data differed by disability status.

**Conclusion**

This study shows a strong association between receiving disability benefits and reduced chances of having children. Moreover, men with disabilities were at reduced chances of having children compared to women with disabilities and to men and women without disabilities. Our findings support policies and programmes aimed at promoting good health during reproductive age as this could reduce early exit from work and early use of disability benefits. Further research is needed to clarify the association between use of disability benefits and all factors related to reproductive health, marriage, and parenthood and family relationships among people with disabilities. Such evidence will help inform both public health and family policies.

**List Of Abbreviations**

CI: - Confidence interval

LISA - Longitudinal Integration Database for Health Insurance and Labour Market Studies

SIMSAM - Swedish Initiative for Research on Microdata in Social and Medical Science

**Declarations**

**Ethics approval and consent to participate**

The Regional Ethical Vetting Board approved all research based on data from the Umeå SIMSAM Lab, including the present paper (Dnr 2010-157-31 Ö). Statistics Sweden made the data anonymous before making it available for this study, therefore obtaining the informed consent for each individual was neither possible nor necessary with this data.

**Consent to publication**

Not applicable

**Availability of data and materials**

The datasets supporting the conclusions of this article are available in the Statistics Sweden repository and can be made available to researchers in accordance with the ethical and legal restrictions regarding Swedish Public Access to Information and Secrecy Act data. Statistics Sweden is a Swedish government agency responsible for official statistics in Sweden. To request access to this data contact: information@scb.se, +46104795000. Researchers requesting access to this data should also have obtained ethical approval from the Swedish Central Ethical Review Board, its contact information is: registrator@etikprovning.se, telephone: +46104750800.
Competing interests

No known competing interest

Funding

The salaries for FN, EL and LV's to contribute to this study is financed by a project headed by Lotta Vikström, that has received funding from the European Research Council (ERC) under the European Union's Horizon 2020 research and innovation programme, Grant Agreement No. 647125, 'DISLIFE Liveable Disabilities: Life courses and opportunity structures across time', 2016–2021. The funders had no role in the study design, data collection, data analysis, data interpretation, or writing of the report.

Authors' contributions

All authors participated in developing the research idea, interpreting the research findings, and the revision of the manuscript. All authors have read and approved the manuscript.

FN: Was responsible for writing the manuscript

ER and FN: Performed the analysis

LV: Secured project funding and supervised the research project

Acknowledgments

The Umeå SIMSAM Lab data infrastructure used in this study was developed with support from the Swedish Research Council (grant number 2008-74919) and with strategic funds from Umeå University. Wallenberg Foundation (Stiftelsen Marcus och Amalia Wallenbergs Minnesfond, MAW 2019.0003).

References

1. Hayford SR: The evolution of fertility expectations over the life course. Demography 2009, 46(4):765-783.
2. Bulatao RA: Values and disvalues of children in successive childbearing decisions. Demography 1981, 18(1):1-25.
3. Kendig H, Dykstra PA, van Gaalen RI, Melkas T: Health of Aging Parents and Childless Individuals. Journal of Family Issues 2007, 28(11):1457-1486.
4. Kelly WR, Cutright P: Economic and other determinants of annual change in U.S. fertility: 1917–1976. Social Science Research 1984, 13(3):250-267.
5. Sandström G, Marklund E: A prelude to the dual provider family – The changing role of female labor force participation and occupational field on fertility outcomes during the baby boom in Sweden 1900–60. The History of the Family 2019, 24(1):149-173.
6. Bumpass LL: What's Happening to the Family? Interactions Between Demographic and Institutional Change. Demography 1990, 27(4):483-498.
7. Ferraro D, Simone AM, Adani G, Vitetta F, Mauri C, Strumia S, Senesi C, Curti E, Baldi E, Santangelo M et al: Definitive childlessness in women with multiple sclerosis: a multicenter study. Neurological Sciences 2017, 38(8):1453-1459.

8. Rahman MHU, Singh A: Socio-economic disparity in the occurrence of disability among older adults in six low and middle income countries. Int J Hum Rights Hea 2019, 12(1):60-75.

9. Grönvik L: Defining disability: effects of disability concepts on research outcomes. International Journal of Social Research Methodology 2009, 12(1):1-18.

10. Stattin M: Retirement on grounds of ill health. Occup Environ Med 2005, 62(2):135-140.

11. Jönsson L, Palme M, Svensson I: Disability Insurance, Population Health and Employment in Sweden. In: Social Security Programs and Retirement around the World: Historical Trends in Mortality and Health, Employment, and Disability Insurance Participation and Reforms. edn. Edited by David A, Wise. Chicago: University of Chicago Press; 2012: 79-126.

12. OECD: Sickness, disability and work: Breaking the barriers. Will the recent reforms make it? Retrieved from http://www.oecd.org/sweden/sicknessdisabilityandworkbreakingthebarriresswedenwilltherecentreformsmakeit.htm 2009a.

13. Oppenheimer VK: Women's Rising Employment and the Future of the Family in Industrial Societies. Population and Development Review 1994, 20(2):293-342.

14. Oppenheimer VK: Cohabiting and marriage during young men's career-development process. Demography 2003, 40(1):127-149.

15. Tumin D: Marriage trends among Americans with childhood-onset disabilities, 1997–2013. Disability and Health Journal 2016, 9(4):713-718.

16. Savage A, McConnell D: The marital status of disabled women in Canada: a population-based analysis. Scandinavian Journal of Disability Research 2016, 18(4):295-303.

17. Namatovu F, Häggström Lundevealler E, Vikström L: The impact of disability on partnership formation in Sweden during 1990-2009. The History of the Family 2020, 25(2):230-245.

18. Sandström G, Namatovu F, Ineland J, Larsson D, Ng N, Stattin M: The Persistence of High Levels of Living Alone Among Adults with Disabilities in Sweden, 1993–2011. Population Research and Policy Review 2020.

19. Simkiss DE, Blackburn CM, Mukoro FO, Read JM, Spencer NJ: Childhood disability and socio-economic circumstances in low and middle income countries: systematic review. Bmc Pediatr 2011, 11.

20. Clarke H, McKay S: Disability, partnership and parenting. Disabil Soc 2014, 29(4):543-555.

21. Sanford C, Newman L, Wagner M, Cameto R, Knokey A-M, Shaver D: The Post-High School Outcomes of Young Adults With Disabilities up to 6 Years After High School: Key Findings From the National Longitudinal Transition Study-2 (NLTS2) (NCSER 2011-3004). In. Edited by International S. Menlo Park, CA: SRI International; 2011.

22. Crawford D, Ostrove JM: Representations of Disability and the Interpersonal Relationships of Women with Disabilities. Women & Therapy 2003, 26(3-4):179-194.

23. Tarasoff LA: Experiences of Women With Physical Disabilities During the Perinatal Period: A Review of the Literature and Recommendations to Improve Care. Health Care for Women International 2015, 36(1):88-107.

24. Parchomiuk M: Social Context of Disabled Parenting. Sexuality and Disability 2014, 32(2):231-242.

25. Nguyen A, Liamputtong P, Horey D: Reproductive Health Care Experiences of People with Physical Disabilities in Vietnam. Sexuality and Disability 2019, 37(3):383-400.

26. Becker H, Stuifbergen A, Tinkle M: Reproductive health care experiences of women with physical disabilities: a qualitative study. Arch Phys Med Rehabil 1997, 78(12 Suppl 5):S26-33.
27. Wickstrom M, Hoglund B, Larsson M, Lundgren M: Increased risk for mental illness, injuries, and violence in children born to mothers with intellectual disability: A register study in Sweden during 1999-2012. Child abuse & neglect 2017, 65:124-131.

28. Lindblad I, Billstedt E, Gillberg C, Fernell E: A register study of life events in young adults born to mothers with mild intellectual disability. Journal of intellectual disabilities : JOID 2014, 18(4):351-363.

29. Lindgren U, Nilsson K, de Luna X, Ivarsson A: Data Resource Profile: Swedish Microdata Research from Childhood into Lifelong Health and Welfare (Umea SIMSAM Lab). Int J Epidemiol 2016, 45(4):1075-1075g.

30. Oppenheimer VK: A Theory of Marriage Timing. American Journal of Sociology 1988, 94(3):563-591.

31. Oláh LS, Bernhardt E: Sweden: Combining childbearing and gender equality. Demographic Research 2008, S7(28):1105-1144.

32. Ronsen M, Sundstrom M: Maternal employment in Scandinavia: A comparison of the after-birth employment activity of Norwegian and Swedish women. J Popul Econ 1996, 9(3):267-285.

Supplementary Tables

Table S1. Odds of later having a first child for men that receive a disability benefit at 20-22 years of age in Sweden. Interaction between disability benefit and marriage.

| Variable                          | Model 1 | Model 2 | Model 3 | Model 5 |
|----------------------------------|---------|---------|---------|---------|
| OR (P-value)                     | 1.0     | 1.0     | 1.0     | 1.0     |
| Disability benefits: No          |         |         |         |         |
| at 20–22yrs                      | 0.04 (0.03–0.04) | 0.04 (0.03–0.04) | 0.04 (0.03–0.04) | 0.02 (0.01–0.03) |
| Year of birth                    |         |         |         |         |
| 1968                             | 1.0     | 1.0     | 1.0     | 1.0     |
| 1969                             | 0.96 (0.92–0.99) | 0.97 (0.93–1.00) | 0.97 (0.94–1.00) |         |
| 1970                             | 0.95 (0.92–0.99) | 0.93 (0.90–0.96) | 0.93 (0.90–0.96) |         |
| Married                          |         |         |         |         |
| No                               | 1.0     | 1.0     |         |         |
| Yes                              | 8.70 (8.45–8.97) | 8.66 (8.39–8.92) |         |         |
| Disability*Married               |         |         |         |         |
|                                 | 4.91 (2.92–8.39) |         |         |         |

COR=Crude Odds Ratio; AOR=Adjusted Odds Ratio; CI=Confidence Interval

Table S2. Odds of later having a first child for women that receive a disability benefit at 20-22 years of age in Sweden. Interaction between disability benefit and marriage.
Table S3. Hazard results of later having a first child for men that receive a disability benefit at 20-22 years of age in Sweden. Interaction between disability benefit and marriage.

| Variable                      | Model 1          | Model 2          | Model 3          | Model 5          |
|-------------------------------|------------------|------------------|------------------|------------------|
| Disability benefits: No       | COR (95 % CI)    | AOR (95 % CI)    | AOR (95 % CI)    | AOR (95 % CI)    |
| at 20–22yrs                   | 1.0              | 1.0              | 1.0              | 1.0              |
|                               | 0.02 (0.02–0.03) | 0.04 (0.02–0.03) | 0.04 (0.03–0.05) | 0.04 (0.03–0.05) |
| Year of birth                 | 1968             | 1.0              | 1.0              | 1.0              |
|                               | 1969             | 0.97 (0.94–0.99) | 0.97 (0.94–1.99) | 0.97 (0.93–1.01) |
|                               | 1970             | 0.91 (0.88–0.94) | 0.91 (0.88–0.94) | 0.97 (0.93–1.00) |
| Married                       | No               | 1.0              |                  |                  |
|                               | Yes              | 5.86 (5.66–6.08) |                  |                  |
| Disability*Married            | 1.0              | 3.01 (2.03–4.51) |                  |                  |

CHR=Crude Hazard Ratio; AOR=Adjusted Hazard Ratio; CI=Confidence Interval
Table S4. Hazard results of later having a first child for women that receive a disability benefit at 20-22 years of age in Sweden. Interaction between disability benefit and marriage.

| Variable                      | Model 1          | Model 2          | Model 3          | Model 5          |
|-------------------------------|------------------|------------------|------------------|------------------|
|                               | CHR (95 % CI)    | AHR (95 % CI)    | AHR (95 % CI)    | AHR (95 % CI)    |
| Disability benefits: No       | 1.00             | 1.00             | 1.00             | 1.00             |
| at 20–22yrs                   | 0.11 (0.09–0.13) | 0.11 (0.09–0.13) | 0.15 (0.13–0.17) | 0.08 (0.06–0.09) |
| Year of birth                 |                  |                  |                  |                  |
| 1968                          | 1.00             | 1.00             | 1.00             | 1.00             |
| 1969                          | 0.96 (0.95–0.98) | 0.97 (0.96–0.98) | 0.97 (0.96–0.98) |                  |
| 1970                          | 0.95 (0.93–0.96) | 0.95 (0.94–0.97) | 0.95 (0.94–0.97) |                  |
| Married                       |                  |                  |                  |                  |
| No                            | 1.00             | 1.00             | 1.00             |                  |
| Yes                           | 2.05 (2.02–2.07) | 2.04 (2.02–2.07) |                  |                  |
| Disability*Married            |                  |                  |                  | 1.00             |
|                               |                  |                  |                  | 6.14 (4.61–8.17) |

CHR=Crude Hazard Ratio; AOR=Adjusted Hazard Ratio; CI=Confidence Interval

**Figures**

**Figure 1**

The Kaplan-Meier survival plot for age at first child