Dissociable processing of emotional and neutral body movements revealed by μ-alpha and beta rhythms

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Abstract

Both when actions are executed and observed, electroencephalography (EEG) has shown reduced alpha-band (8–12 Hz) oscillations over sensorimotor cortex. This ‘μ-alpha’ suppression is thought to reflect mental simulation of action, which has been argued to support internal representation of others’ emotional states. Despite the proposed role of simulation in emotion perception, little is known about the effect of emotional content on μ-suppression. We recorded high-density EEG while participants viewed point-light displays of emotional vs neutral body movements in ‘coherent’ biologically plausible and ‘scrambled’ configurations. Although coherent relative to scrambled stimuli elicited μ-alpha suppression, the comparison of emotional and neutral movement, controlling for basic visual input, revealed suppression effects in both alpha and beta bands. Whereas alpha-band activity reflected reduced power for emotional stimuli in central and occipital sensors, beta power at frontocentral sites was driven by enhancement for neutral relative to emotional actions. A median-split by autism-spectrum quotient score revealed weaker μ-alpha suppression and beta enhancement in participants with autistic tendencies, suggesting that sensorimotor simulation may be differentially engaged depending on social capabilities. Consistent with theories of embodied emotion, these data support a link between simulation and social perception while more firmly connecting emotional processing to the activity of sensorimotor systems.

Key words: action simulation; emotional body movement; time-frequency EEG; mu suppression; beta enhancement

Introduction

We can rapidly infer emotions from nonverbal face and body cues. Although this process could arise from high-level theorizing and reasoning, a growing body of evidence suggests that we understand others through our own sensorimotor systems, specifically by internally simulating their external movements (Niedenthal, 2007; Barsalou, 2008; Wood et al., 2016). Consistent with this idea, research suggests that body movements provide a rich and immediate source of social information (Beall et al., 2008) that can be extracted rapidly and automatically (de Gelder, 2006).

Behavioral evidence for simulation of emotion comes from emotional mimicry, that is, quick and spontaneous matching of another’s expressions (Dimberg and Thunberg, 1998). More than a simple motor matching process, mimicry appears to arise from mental simulations of others’ emotions (Moody et al., 2007; Wood et al., 2016) and has been shown to be impaired in conditions associated with social processing deficits such as autism spectrum disorder (ASD) (McIntosh et al., 2006; Moody and McIntosh, 2006; Beall et al., 2008; Oberman et al., 2009). Likewise, emotional contagion, in which individuals ‘catch’ others’ emotional states automatically and unconsciously,
been associated with facial, vocal and postural mimicry (Hatfield et al., 1993).

In interpreting such 'embodied emotion' phenomena, researchers have drawn on an extensive literature documenting neural simulation within sensorimotor systems during observation of others' actions (Decety, 1996; Gallee and Goldman, 1998; Jeannerod, 2001, 2006). Both when actions are actually executed and when they are merely imagined or observed, there are visible decreases in the power, or magnitude, of oscillatory neural activity over sensorimotor cortex in electroencephalography (EEG; for a review, see Fox et al., 2016). Known as the μ-rhythm, this oscillatory activity is typically measured over central sensors in the alpha (8–12 Hz) frequency band (Niedermeyer and da Silva, 2005). Paralleling the report of mirror neurons (Rizzolatti et al., 2001), which respond to both observed and executed actions, μ-suppression has been argued to reflect the active engagement of sensorimotor areas during action simulation (Niedermeyer and da Silva, 2005; Ulloa and Pineda, 2007).

Differential μ-suppression has been associated with action understanding (Rizzolatti et al., 2001; Ulloa and Pineda, 2007), empathy and social processes (Kilner et al., 2006; Pineda and Hecht, 2009) and social processing impairments in ASD (Oberman et al., 2005; Oberman and Ramachandran, 2007). Yet, despite these claims, the link between μ-suppression, action simulation and the role of sensorimotor systems in perceiving others' emotions has not been directly tested. Although previous studies have examined the effect of emotion on μ-suppression (Perry et al., 2010; Moore et al., 2012; Moore and Franz, 2017), the stimuli used in these experiments are not associated with body movements and/or are unlikely to directly initiate mental reenactment: for example, static faces (Moore et al., 2012; Moore and Franz, 2017) or situational information about the painfulness of a stimulus applied to another individual (Perry et al., 2010).

In contrast, body movements form an ideal stimulus category to test whether action simulation plays a role in perceiving others' emotions. Body movements can be rapidly perceived and categorized even from relatively sparse visual information, as in point-light displays (PLDs; Blake and Shiffrar, 2007). By ‘scrambling’ the global stimulus structure of the PLDs, we can pre-process scrambled stimuli, in line with previous studies (Ulloa and Pineda, 2007). Additionally, many studies limit their measurements to the alpha band, potentially ignoring sensorimotor activity in other frequency ranges (Hobson and Bishop, 2016). In particular, activity in the central beta band (16–20 Hz) is suppressed by imagined, observed and executed movement, similar to μ-suppression (McFarland et al., 2000; Babiloni et al., 2002; Zaeppfli et al., 2013). Beta activity has also been reported to increase during the first 500 ms of viewing emotional stimuli such as angry faces (Guntekin and Basar, 2014).

To address these issues, we measured EEG power spectra over 128 electrode locations in a dense whole-head electrode array across a broad range of frequencies between 4 and 20 Hz. Directly comparing biologically plausible coherent vs scrambled PLDs allowed us to more rigorously control for brain activity related to low-level visual stimulation. By implementing these more rigorous standards, our study provides a novel and meticulous approach to the link between μ-suppression, action simulation and social perception.

Methods
Participants
Forty-six undergraduate participants were recruited for either monetary compensation or partial course credit. Eight participants were excluded from the analyses for the following reasons: (i) technical issues with EEG recording (n = 5); (ii) failure to comply with experiment instructions (n = 1); and (iii) performance 2.5 s.d. lower than the average d-prime (d′) score (n = 2). Thus, a total of 38 participants (ages 18–23; 16 females) were submitted for data analysis. This final sample size complied with recent recommendations for adequate power to measure μ-suppression (Hobson and Bishop, 2017). All experimental procedures were reviewed and approved by Claremont McKenna College’s Institutional Review Board, and informed consent was obtained in writing from all participants prior to the experiment.

Stimuli
Stimuli consisted of whole-body actions depicted in PLD videos, which control for irrelevant visual body shape information (Figure 1A). The PLDs were 3 s video clips used in Atkinson et al. (2012), adapted from a larger database developed by Atkinson et al. (2004).

PLD videos consisted of 11 different actors of both genders (5 males, 6 females) portraying whole-body expressions of different emotions or of emotionally neutral common everyday actions. The emotional content condition contained PLDs of...
PLDs. Each action coherency condition included 18 distinct emotional and 6 distinct neutral PLDs.

Procedure

During the experiment, participants viewed two separate blocks of coherent and scrambled biological motion stimuli (Figure 1B) while EEG recordings were taken. The order of the two blocks was counterbalanced across participants. Each individual trial consisted of a single, 3 s PLD, followed by an intertrial interval (ITI) during which the last frame was presented as a static image for another 2 s. To ensure attentive viewing, participants performed a continuous one-back task in which they monitored for the immediate repetition of each video (i.e. the same configuration and movement of the dots, presented twice in a row); participants were explicitly instructed to look for an exact repetition of the dot configurations and their local motion, rather than just the general gist. Given the dynamic nature of the PLDs, the duration of each stimulus and the high number of non-immediate stimulus repetitions across the experiment, this task was substantially more challenging than a traditional one-back paradigm and required participants to pay attention throughout the experiment.

In each block, the number of presentations of the emotional and neutral PLDs were equated by presenting each of 18 emotional PLDs 3 times and each of 6 neutral PLDs 9 times, respectively, for a total of 54 trials per content type (emotional vs neutral). Thus, each action coherency block (coherent vs scrambled) consisted of 108 (2 × 54) trials for each content type. Across the two experimental blocks, PLDs were repeated following 12 randomly selected trials within each block to create one-back trials for the task, generating a total of 240 trials (216 stimulus presentations plus 24 one-back trials). Aside from the 24 one-back trials, all stimuli within each action coherency block were pseudo-randomly interleaved to ensure that there were no consecutive repetitions of each video. Participants completed 10 practice trials (5 coherent and 5 scrambled, each containing a one-back trial) before beginning the main experiment.

After EEG data acquisition, participants answered two multiple-choice questions where they were asked to select all emotions and actions that they thought they saw during the experiment, out of six word choices for emotion (sadness, happiness, anger, fear, disgust and surprise) and action (touching toes, jumping on one foot, walking, jumping jacks, kicking and kneeling down). They then completed an online version of the autism quotient questionnaire (Baron-Cohen et al., 2001).

EEG data acquisition and analysis

EEG data was collected using a BioSemi ActiveTwo system (Biosemi B.V., Amsterdam, the Netherlands) with 128 channels of active electrodes inserted in fitted headcaps. Two additional electrodes were adhered bilaterally to participants’ mastoids to serve as reference channels for data import. Recorded signals were digitized continuously at 512 Hz with a hardware low-pass at one-fifth of the sampling rate. Before beginning data collection, electrodes were adjusted until offsets were between −20 and 20 μV and there were no obvious slow drifts in the online data output.

Data preprocessing was performed offline using the EEGLAB toolbox (Delorme and Makeig, 2004) for MATLAB (Mathworks, Natick, MA). Following import into EEGLAB, data were resampled at 500 Hz. Linear detrending was conducted on
continuous data to remove direct current offsets. A high-pass filter of 0.5 Hz was applied to remove slow voltage drifts via a two-way least-squares finite impulse response filter. Data were not re-referenced and no low-pass filters were applied in order to avoid distorting the power spectra (Luck, 2014). For each trial, 3600 ms epochs were extracted around the 3000 ms stimulus, −400 ms pre- to 3200 ms post-stimulus onset. Data were baseline-corrected to the pre-stimulus period to correct for initial electrode offsets.

All one-back trials and other trials where participants made a motor response were removed from data analyses to exclude potential motor preparatory activity. To identify and remove EEG artifacts (e.g. head, eye and jaw movements as well as electrical and sensor noise), we conducted an independent components analysis on the remaining trials for each participant using the second-order blind identification method (Belouchrani et al., 1997; Tang et al., 2005). The remaining task-related components were projected back onto the scalp (Jung et al., 2000).

Event-related potentials (ERPs) for each condition were compared using threshold-free cluster enhancement (Mensen and Khatami, 2013), with default parameters of extent of 0.666 and height of 1. Grand average waveforms for each subject and condition were entered into a 2 × 2 analysis of variance (ANOVA) with content (emotional/neutral) and action coherency (coherent/scrambled) as factors. Resulting clusters of significant activity were corrected for multiple comparisons using a permutation approach (2500 permutations).

Time-frequency analysis was performed using the FieldTrip toolbox (Oostenveld et al., 2011) for MATLAB. For each participant and condition (Emotionalcoherent, Emotionalscrambled, Neutralcoherent and Neutralscrambled), power was computed averaging across trials at each electrode location using a Morlet wavelet (width = 7) for a conservatively large frequency range of 4–20 Hz and an analysis window centered on −0.2 to 3 s, sliding in steps of 0.01 s. Time-frequency data was log_{10} transformed to normalize the frequency distribution. We compared power across conditions using a dependent sample two-tailed t-test with a nonparametric cluster-based Monte Carlo permutation test (1000 repetitions) to correct for multiple comparisons. Significance was assessed using an overall threshold of P = 0.05, with significance of P = 0.025 at each tail. Effects before 1 s and beyond 2 s from stimulus onset were excluded as a conservative measure due to potential confounds in alpha-band activity associated with stimulus onset and offset (Fox et al., 2016; Hobson and Bishop, 2017).

Results

Behavioral results

As a manipulation check following EEG recording, we asked participants to indicate which emotions and actions they had seen during the task. Participants were highly accurate at identifying the emotions portrayed by the PLDs, with 94.7% (36/38) of participants correctly selecting at least two of the emotions and 89.5% (34/38) selecting all three emotions. Identification rates for the neutral actions were similarly high, with 97.4% (37/38) of participants correctly selecting at least two of the three actions and 81.6% (31/38) selecting all three actions. A further question is whether viewing emotional movements influences subjective reports of emotional states. However, we could not examine this in the current design because different emotional movements were pseudo-randomly interleaved along with neutral movements, making it unclear what the cumulative effect would be. To assess participants’ monitoring performance during the task, we computed d’ scores (d’ = ZFalse Alarm Rate – ZHit Rate) (Macmillan and Creelman, 2004) for the one-back task (mean = 3.25, s.d. = 0.65). A 2 × 2 repeated-measures ANOVA with factors of content (emotional/neutral) and action coherency (coherent/scrambled) found no significant main effects of content \( F(1,37) = 0.06, P = 0.81, \eta^2_p = 0.002 \) or action coherency \( F(1,37) = 0.94, P = 0.34, \eta^2_p = 0.03 \), though the interaction of content × action coherency approached significance \( F(1,37) = 2.92, P = 0.10, \eta^2_p = 0.07 \), driven by a larger difference in accuracy for emotional content in the coherent vs scrambled conditions. However, the overall lack of significant differences between content and action coherency levels suggests that participants were deploying attention across all conditions.

Event-related potentials

To assess whether our experimental conditions evoked differential neural responses, as well as assessing the quality of the neural data, we examined the ERP response. Grand average data time-locked to the onset of the stimulus were entered into a 2 × 2 ANOVA with content (emotional/neutral) and action coherency (coherent/scrambled) as factors using a threshold-free cluster enhancement approach (Mensen and Khatami, 2013). Taking advantage of spatial and temporal relationships in ERP data, this analysis has been demonstrated to have high sensitivity to detect meaningful signals while maintaining statistical integrity (Mensen and Khatami, 2013; Pernet et al., 2015). Figure 2 displays the temporal distribution of significant main effects for each factor, along with representative scalp plots and grand average waveforms. The interaction effect failed to reach significance in any sensors or time windows (all \( P > 0.2 \)).

A significant effect of action coherency (Figure 2A) emerged relatively early in the trial, from 288 to 330 ms after stimulus onset in left frontal and anterior temporal sensors (Figure 2A, left). Reflecting greater activity to coherent vs scrambled PLDs (Figure 2A, top right), the timing of this effect is largely consistent with latencies reported in prior ERP studies (for a review, see Thompson and Parasuraman, 2012). However, no significant effects were observed across the rest of the trial (Figure 2A, bottom right). Therefore, to the extent that coherent biological motion stimuli elicit additional perceptual and cognitive processing, these effects are unlikely to explain any observed differences in µ-suppression during the time window of interest (1–2 s post-stimulus onset).

In contrast, significant neural activity associated with content was observed across the time course of action perception (Figure 2B). Although the largest effects occurred −2 s after stimulus onset, significant clusters were visible within the time window of interest for µ-suppression effects, with a local peak significance of ~1.67 s after stimulus onset (Figure 2B, bottom right). Examination of the evoked signals for each condition within this time window revealed this effect to be driven largely by the response to coherent emotional PLDs (Figure 2B, top right). Although the timing of this effect is similar to a previously reported slow-wave potential over centrotemporal sensors that is associated with viewing of emotional pictures (Cuthbert et al., 2000), the response observed here differs notably both in terms of scalp topography (Figure 2B, left) and level of sustained activity, complicating the interpretation of this result. Nonetheless, the divergence of this ERP scalp topography from the canonical distribution of µ-suppression over central sensors suggests that cognitive processing indexed by this component may be separable from action simulation.

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**Time-frequency analysis: coherent vs scrambled action**

Our first time-frequency analysis examined whether the comparison of coherent vs scrambled PLDs elicited significant μ-suppression over central sensors, as reported in previous studies. Averaging across power spectra for emotional and neutral stimuli, we then subtracted the 4-20 Hz power spectra for scrambled PLDs from the power spectra for coherent PLDs. If biologically plausible movements produce additional sensorimotor activity compared to non-biological motion, as posited by simulation theory, we should see significant reductions in power over central sensors within the alpha and/or beta frequency bands.

**Figure 3** displays the difference in average power for coherent vs scrambled PLDs across all 128 channels. Consistent with previous reports of μ-suppression (e.g. Fox et al., 2016), we observed a clear band of spectral power suppression between ~9 and 12 Hz from roughly 1 to 2 s after stimulus onset (Figure 3A, black box). Statistical analyses of this time and frequency window revealed significant reductions at central and temporal sensors roughly overlying sensorimotor cortex (Figure 3B), surviving cluster-based multiple-comparisons permutation correction at $P = 0.03$. Calculation of effect size using Cohen’s $d$ with pooled standard deviation for within-subjects designs (Cohen, 1988; Morris and DeShon, 2002) likewise suggested moderate practical
Main effect of action coherency. (A) Time-frequency plot of average power spectrum from 4 to 20 Hz for coherent–scrambled PLDs across all 128 channels. In the time window between 1 and 2 s, we observed reduced power within the alpha band of 9–12 Hz (black box). (B) Statistical analysis of the highlighted time and frequency window using a paired samples t-test revealed significant suppression ($P < 0.05$, permutation-corrected for multiple comparisons) in centroparietal and temporal sensors, as indicated by Xs.

Time-frequency analysis: emotional vs neutral movement

To assess the main experimental question of whether emotional movements elicit greater μ-suppression, we next compared spectral power for emotional and neutral content conditions after controlling for low-level motion information. Thus, this analysis used a difference of differences, or double subtraction, in which we first subtracted the response to the scrambled PLDs ('s') from the response to the equivalent coherent versions ('c'): \( \text{Emotional}_c - \text{Emotional}_s = \text{Emotional}_\text{coh} - \text{Emotional}_\text{scram} \); \( \text{Neutral}_c - \text{Neutral}_s = \text{Neutral}_\text{coh} - \text{Neutral}_\text{scram} \) and then compared the two \( \text{Emotional}_c - \text{Emotional}_s - \text{Neutral}_c + \text{Neutral}_s \).

As shown in Figure 4A, this analysis identified a more focal time window of suppression in the alpha (Figure 4A, red box) and beta (Figure 4A, orange box) frequency bands, from ~1.3 to 1.5 s after stimulus onset. Examining the specific ranges associated with spectral peaks for μ-alpha and beta, we found differential distribution of significant suppression effects (Figure 4B and C), surviving cluster-corrected multiple-comparisons correction at $P < 0.01$ ($d = -0.6$).

Whereas power reductions in the beta band (16–20 Hz) were found over central and frontal sensors and were most prominent at central locations (Figure 4B), μ-suppression in the canonical 9–12 Hz range was distributed over central and occipital sensors (Figure 4C). Examining the full range of frequencies from 4 to 20 Hz within this time window confirmed these findings, with significant reductions in power (cluster-corrected $P < 0.01$) emerging from ~9 to 20 Hz and progressing from occipital and central sensors to more frontal electrodes with increasing frequency (Figure 4D).

Given the complex nature of the double subtraction comparison, a further question is which conditions are driving the overall effect. Therefore, as a follow-up, we also plotted the simple subtractions of emotional – neutral in the coherent (Figure 5A, left) and scrambled (Figure 5A, right) conditions from 1 to 2 s. As expected, reductions in power were clearly visible within the time window of interest from 1.3 to 1.5 s for coherent stimuli but not for scrambled stimuli. This finding supports the idea that differential μ-suppression to emotional vs neutral stimuli reflects cognitive processes rather than low-level visual differences between the stimuli.

However, this analysis does not specify the direction of these significant suppression effects with respect to the Emotional\(_c\) vs Neutral\(_c\) comparisons. Therefore, in simple subtractions, we directly compared the response to emotional and neutral stimuli while controlling for low-level visual factors by subtracting Emotional\(_\text{coh}\) – Emotional\(_\text{scram}\) (Figure 5B, left). Consistent with previous reports of μ-alpha suppression, coherent vs scrambled emotional body movements were associated with decreased power between 9 and 12 Hz (Figure 5B, left, black box). These effects were strongest over central and occipital sensors (Figure 5C, left), in line with the results of the double subtraction; in contrast, no beta suppression was observed in this contrast. Surprisingly, however, comparison of Neutral\(_\text{coh}\) – Neutral\(_\text{scram}\) revealed enhanced power in the beta range between 1.3 and 1.5 s (Figure 5B, right, dashed black box) and this effect was strongest for central and frontal electrodes (Figure 5C, right). Thus, despite similar underlying sources, μ-suppression effects in the alpha and beta bands may reflect dissociable cognitive processes that are differentially influenced by emotional stimulus content.

Individual differences analysis: median-split by AQ

The above results are consistent with the idea that μ-suppression indexes action simulation, and this simulation is enhanced by emotional content. To help interpret this finding, we examined whether μ-suppression varies with autistic tendencies,
given that ASD has previously been correlated with reduced μ-suppression. Although our sample may be underpowered for individual difference tests, we performed an exploratory analysis based on AQ scores (Baron-Cohen et al., 2001). Because of the relatively small sample size and restricted range of AQ scores within our sample, we expected that power would be an issue for a correlational analysis; therefore, we opted to perform a median-split analysis on the data. Across participants, the mean AQ score was 19.3 (median = 19.5, s.d. = 5.13), within the typical range of social functioning (i.e. 11–21). Dividing participants using a median-split, we obtained two groups with low AQ (mean = 15.4, s.d. = 2.99) and high AQ (mean = 23.3, s.d. = 3.51), associated respectively with neurotypical performance and above-average autistic tendencies.

Comparing the neural responses to emotional actions in these two groups, we observed striking differences in the extent of alpha-band suppression associated with coherent vs scrambled stimuli (Figure 6A). Whereas the low AQ group showed extensive reductions in power across the window from 1 to 2 s (Figure 6A, top), μ-suppression effects were weaker in the high-AQ group (Figure 6A, bottom). Additionally, these effects varied in their scalp distribution, with the low-AQ group showing significant effects across central, parietal and temporal sensors (Figure 6B, top, white dots; $P = 0.002, d = −0.86$). In contrast, alpha-band suppression in the high-AQ group was concentrated over occipital sensors, though the within-group EmotionalCoherent vs EmotionalScrambled contrast did not reach significance at any sensor location (Figure 6B, bottom). To directly assess these group-level differences, we developed an index of μ-alpha suppression by defining sensors of interest (SOIs) using the significant sensors from the initial contrast of coherent – scrambled across both emotional and neutral stimuli, computing the average power between 9 and 12 Hz from 1 to 2 s after stimulus onset. Comparing power in these ‘μ-alpha SOIs’ to that over occipital electrodes within the same time range revealed higher suppression at μ-alpha SOIs in the low-AQ group, despite similar levels of occipital alpha suppression (Figure 6C). Confirming this effect, we found a significant interaction of group by location [$t(36) = −2.18$, $P = 0.036, d = −0.71$]. Again, these effects were dissociable from the enhancement of beta activity for neutral actions (Figure 7), which was strongest in the low-AQ group ($P = 0.0001$, $d = 0.97$), suggesting that oscillations in the alpha and beta bands contribute differentially to action simulation processes.

Discussion

Although μ-suppression is often cited in support of simulation theory, little work has focused on the selective effects of
emotional actions on μ-power. Further, many studies of μ-suppression conflate perceptual, attentional and action simulation processes, making it difficult to disentangle which aspects of emotion might drive reductions in oscillatory power. Here, we more effectively isolated μ-rhythms specific to action simulation processes by implementing recent methodological recommendations, such as using high-density EEG and controlling for perceptual and attentional confounds (Fox et al., 2016; Hobson and Bishop, 2016; Bowman et al., 2017). With these controls, we successfully replicated previous reports of greater μ-alpha suppression for coherent vs scrambled biological motion from PLDs (Ulloa and Pineda, 2007). Further, we identified ERP evidence for early cognitive differences in processing coherent vs scrambled stimuli. However, they occur substantially prior to our window of interest (1–2 s) for μ-alpha suppression effects, suggesting that the sustained process of action-related simulation is likely separable from these more immediate perceptual responses to coherent biological stimuli.

Simultaneously mapping power changes over the whole head for emotional vs neutral stimuli, we found that μ-alpha suppression over central sites was associated with the perception of emotional content, supporting our hypothesis that action-related simulation processes are additionally engaged by emotional content in others’ movements. At the same time, we observed alpha suppression over occipital sensors even after controlling for low-level visual properties, supporting the idea that low-level visual features alone cannot explain the attentional salience of emotional content (Nummenmaa et al., 2006).

Given the inherent salience of emotional stimuli, effects of emotion and attentional processing are inevitably difficult to completely dissociate. However, the lack of performance differences in an attentionally demanding one-back task, along with the markedly different scalp distribution of the ERP wave-
form and μ-alpha suppression, suggest that the time-frequency effects cannot be wholly accounted for by differences in motivated attention associated with emotional stimuli.

Although many studies of μ-suppression focus on the alpha band, we also found significant effects in the beta band, from ~16 to 20 Hz, distributed over frontocentral sensors. Simple subtractions of coherent vs scrambled stimuli revealed effects for emotional stimuli in the canonical μ-alpha (9–12 Hz) range, whereas these beta-band effects appeared to arise instead from enhancement of the response to coherent emotionally neutral actions. In contrast to the extensive literature on μ-suppression, beta enhancement is more difficult to explain. Beta enhancement has been previously reported for inhibition of motor responses in GO-NO GO tasks (Zaepffel et al., 2013). In studies of sensorimotor simulation, beta enhancement or event-related synchronization is often described as a ‘rebound’ effect, because it is seen as a brief increase in power that follows a period of suppression during which motion is executed (Hari et al., 1998; Babiloni et al., 2002). However, beta power does not only passively reflect the lack of movement; rather, it has been shown to be relevant to the active maintenance of the current motor set or cognitive state (Gilbertson et al., 2005; Engel and Fries, 2010). Notably, Brinkman et al. (2014) reported differential roles of alpha vs beta oscillations in suppressing task-irrelevant regions vs inhibiting movement parameter computations.

Consistent with this idea, our finding of differential effects of coherency for emotional and neutral stimuli suggests that beta and μ-alpha are not interchangeable measures of sensorimotor simulation, as often assumed in previous μ-suppression studies (Hobson and Bishop, 2017). Rather, μ-alpha and beta may reflect dissociable cognitive processes that are involved in the perception of different types of body movement. One clue regarding the nature of this distinction comes from recent evidence that μ-alpha reflects tactile rather than motor aspects of mirroring (Coll et al., 2015; Coll et al., 2017). Consistent with this idea, μ-alpha and beta power are inversely related to functional magnetic imaging (fMRI) signal in primary somatosensory and motor cortex, respectively (Ritter et al., 2009). Although the lack of a tactile component to our stimuli complicates any direct comparison between these results, it is interesting to note that our μ-suppression effects appear to be driven by perception of emotional content, which is known to engage somatosensory networks (e.g. Adolphs et al., 2000). In this light, we can hypothesize that μ-suppression in the alpha band represents somatosensory aspects of emotion simulation, whereas beta enhancement maintains action simulation processes while inhibiting motor output.

Yet, if beta enhancement is tied to motor inhibition, why did emotional body movements fail to elicit this processing? One possible answer arises from the evolutionary significance of emotional, as opposed to neutral, body movements. While it is appropriate to inhibit unwanted movement and maintain the status quo in response to neutral stimuli, inhibiting automatic motor responses may be less desirable when confronted with emotional stimuli. Because automatic facial and bodily responses to emotional stimuli can serve social goals (Reichmann-Decker et al., 2009; Moody et al., 2017), it may be disadvantageous for the sensorimotor system to inhibit potentially beneficial automatic approach/avoidance responses to the intensely affective bodily movement of others.

Furthermore, the emotional body movement of others may be more difficult to mimic because emotional expressions are idiosyncratic (Wibbeler et al., 2011). Therefore, there may be less need for an effortful suppression of an alternative motor state. Conversely, common actions such as touching one’s toes are performed in a similar manner by most individuals. As such, neutral actions may more easily activate automatic motor programs which would then have to be inhibited. The increase in beta power may reflect this inhibition of physical movement in order to maintain the current state of mental simulation.

If this mental simulation of others’ body movements allows us to better understand their emotional states, then it follows that individuals who are less sensitive to social cues may exhibit reduced neural correlates of action simulation. To address this question, we conducted an exploratory analysis examining alpha and beta power for two subgroups median-split by AQ, a behavioral measure of emotional intelligence and receptivity to social information. Looking at the neural response to emotional stimuli, we found that whereas the high- and low-AQ groups did not differ in terms of occipital alpha, the population-typical low-AQ group exhibited more pronounced μ-alpha suppression over centroparietal SOIs. Consistent with previous results from clinical populations with ASD (Oberman et al., 2005; Oberman and Ramachandran, 2007; Dumas et al., 2014), these results suggest that autistic tendencies are associated with selectively lower simulation activity for emotional content, even when occipitoparietal alpha-band activity suggests similar levels of attentional engagement.

At the same time, we observed strongest beta enhancement for neutral actions in the low-AQ group. Notably, this effect showed high statistical and practical significance, as indicated by a low P value and high effect size, despite having only half the sample size of the original experiment. Thus, rather than reflecting variability in simulation processes associated with autistic tendencies, these results confirm that, like μ-alpha, the beta effects are driven by ‘normal’ receptiveness to social information from the movements of others.

However, it is important to note that the AQ is not a clinical measure, but rather a preliminary tool for identifying traits associated with ASD in adults with normal intelligence (Baron-
Cohen et al., 2001). Within our undergraduate sample, only one participant (AQ = 34) had an AQ score above the cut off indicating high likelihood of diagnosable ASD. Therefore, it is possible we would observe more pronounced differences in μ-alpha and beta effects in clinically diagnosed ASD populations.

Collectively, these results demonstrate that emotional content in body movements elicits higher levels of alpha-band μ-suppression, whereas increased power in the beta range reflects dissociable responses to neutral, rather than emotional, coherent actions. Consistent with theories of embodied emotion, these data support a link between simulation and social perception while more firmly connecting emotional processing to the activity of sensorimotor systems. In line with recent data identifying dissociations between sensorimotor processes in the alpha and beta bands, future research should further investigate the computational implications of separable alpha suppression and beta enhancement effects during action simulation.

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