Examining the Association between Polish Migrant Status and Health Preferences Using a Novel Application of a Smaller Design EQ-5D-5L Valuation Study

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Abstract

Introduction Migrants have different utilisation of healthcare services and health-related behaviours than host populations. A potential factor that may contribute to the notable differences in healthcare use and health-related behaviours between migrants and host populations is how these groups value health. Those who place a high value on health have greater healthcare-seeking practices than those who do not.

Objective The aim of this study was to examine how Polish migrants and native Irish differ in health state utility valuations using a novel application of a smaller design EQ-5D-5L valuation study.

Methods This study uses health preferences as a predictor of how one values health. We examined the EQ-5D-5L health preferences of 119 Polish migrants and 123 native Irish, both residing full-time in Ireland. To do so, we used a novel application of a smaller design EQ-5D-5L valuation study that consisted of 30 health states and a targeted sampling strategy coupled with a Bayesian statistical nonparametric model. We collected data from June 2018 to September 2019.

Results Our results highlight that Polish migrants and native Irish differ in their health preferences for and valuation of severe health states. Polish migrants place meaningfully higher utility valuations of 0.1 or more on the three most severe health states compared with the native Irish.

Conclusion This study can provide an understanding of a potential new factor underpinning some of the disparities in healthcare utilisation and health-related behaviours among migrants and host populations in Europe. This study also provides proof of principle for using a smaller design EQ-5D-5L valuation study to explore differences in health preferences among other minority subgroups, which can otherwise be hard to uncover when using the secondary analysis of national EQ-5D-5L valuation studies.

Key Points for Decision Makers

Polish migrants apply a higher utility valuation to more severe health states than native Irish respondents.

Differences in health preferences potentially infer differences in health-related behaviours between migrants and natives.

Smaller design EQ-5-D-5L valuation studies can be used in examining the health preferences of hard-to-research subgroups.

1 Introduction

European countries have become increasingly popular destinations for migrants from other European states as well as migrants from outside of Europe. The European Union (EU) has a diverse stock of migrant populations. Migrants make up an estimated 13.2% of the EU’s total population, with individuals born in another EU member state totalling 5.5% of the population and those born in a non-EU country making up 7.7% of the population [1]. Numerous studies have examined the healthcare service use and health-related behaviours of migrants relative to host populations in Europe [2–4]. In Europe, migrants are noted to have different utilisation of healthcare services than host populations, with the differences most evident in preventive healthcare uptake, as migrants consistently have lower uptake than host populations [5–7]. Migrants are also shown to have different health-related behaviours...
that host populations, and have been noted to engage in more unhealthy behaviours when compared with host populations in Europe, such as increased smoking and alcohol consumption [8].

Marked differences and inequalities in the healthcare utilisation and health-related behaviours of migrants and host populations in Europe have been identified in the literature. Understanding the many factors that drive these disparities is pivotal for designing, adapting, and implementing a migrant health policy in Europe, and a multitude of studies have examined such factors, such as the formal and informal barriers to healthcare faced by migrants, transnationalism and cultural attitudes and beliefs regarding health and healthcare held by migrants [9–13]. A potential factor that may contribute to some of the notable differences in the healthcare use and health-related behaviours of migrants compared with host populations is how these groups value health.

An individual’s valuation of health has been shown to help determine their use of healthcare and health-related behaviours, as individuals who place a high value on health have greater healthcare-seeking practices, especially preventive care use, than those who do not [14–16]. Determining how migrants and host populations differ in their valuation of health is an under-researched area despite the potential insight this may provide into the noted differences between groups in their healthcare use and health-related behaviours. Health preferences are a predictor of an individual’s valuation of health. Health preferences as a variable, proxies the value an individual places on health (such as how an individual values different health state utilities), which can likely inform an individual’s healthcare use and health-related behaviours [17, 18].

In this study, we examine the association between Polish migrant status and health preferences as measured by the EQ-5D-5L health descriptive system. More specifically, we examine the health preferences of Polish migrants residing full-time in Ireland relative to a group of native Irish also residing full-time in Ireland. The Polish migrant group are the largest migrant group residing in Ireland [19]. As such, understanding how Polish migrants and native Irish differ in their health preferences, and ultimately in their valuation of health, may provide new insight into the reasoning that underpins some of the disparities in healthcare use and health-related behaviours between migrants and host populations in Europe.

## 2 Methods

The usual manner in which to conduct any subgroup analysis of health preferences, such as comparing migrants and a host population, is to use a secondary analysis of national EQ-5D-5L valuation studies. Numerous countries have carried out national valuation studies [20–23], and other valuation studies have emerged that use novel study design methods to examine EQ-5D-5L health preferences [24]. National valuation studies not only primarily provide researchers with data on health preferences to inform cost-utility analysis (CUA) but they also provide researchers with a potential opportunity to compare the health preferences of various subgroups. However, one issue is that national valuation studies are not primarily designed to examine differences in health preferences between subgroups and may not be powered to detect such differences.

When using a secondary analysis of national valuation studies to compare subgroup health preferences, an issue can be the relative sample size of the subgroups uncovered in the overall sampling of the general public [17, 18, 25–27]. Solutions to this issue have included oversampling of specific subgroups alongside national valuation studies, or, as recommended by Sculpher and Gafni [28], conducting an additional survey to uncover subgroups in greater detail. The authors suggest this additional survey can complement national valuation studies and would allow for the targeted sampling of subgroups. This targeted approach requires a novel application of EQ-5D-5L valuation exercises whereby a reduced sample size values a smaller number of health states compared with national valuation studies [29, 30]. Ultimately, this requires a smaller design valuation study, which can be deployed to target subgroups for whom preferences are thought to differ, such as migrants and host populations.

### 2.1 Survey Design and Sampling Strategy

To capture data on health preferences, we estimated preferences for the dimensions of the EQ-5D-5L using the composite time trade-off (cTTO) approach [31]. cTTO data were collected for a group of Polish migrants and native Irish, both residing full-time in Ireland. Polish migrants were selected for the analysis as they are the largest non-Irish population living in Ireland, with over 120,000 Polish migrants residing in Ireland [19]. The majority of Polish migrants in Ireland are aged between 30 and 40 years, making up 47% of the Polish population in Ireland, with 50.49% males and 49.51 females [19]. Each respondent valued six practice states plus one block of 11 cTTO states (giving 11 observations per respondent), which all respondents completed in English. A sample size of 240 (120 Polish migrants and 120 native Irish) was deemed appropriate as it was the maximum that could be collected based on the resources available for this research. Data were collected from June 2018 to September 2019, with the EuroQol Research Foundation providing the orthogonal design of health states for valuation. All interviews were collected using the EuroQol Research Foundation’s
2.2 Experimental Design

We used a similar orthogonal design as Yang et al. [30], which, in this study, entailed 25 health states, including the 55555 'pits' state plus five mild health states to compensate for an underrepresentation of mild states, for a total of 30 EQ-5D-5L health states to be valued. Using an orthogonal design allows us to value a smaller number of health states and use a smaller sample size compared with that of a full national valuation study, which uses 86 cTTO health states and a minimum of 1000 respondents [22]. Yang et al. [29, 30] have examined and compared estimates of the smaller design EQ-5D-5L valuation study with the traditional EQ-5D-5L valuation study of 86 health states. Both papers note that the smaller design performs well compared with the larger design and does not lead to significant increases in prediction errors when modelling cTTO data.

2.3 Blocking

The 30 health states were divided into three blocks of 11 states using the blocking algorithm included in the “AlgDesign” package in R. The blocking algorithm divides the states over the blocks to ensure that within-block variance is maximised so that observations on the full utility range is achieved.

2.4 Interviewers and Respondents

We collected cTTO state valuations for the three blocks of health states for Polish migrants who were born in Poland and residing full-time in Ireland and a group of native Irish who were born in Ireland and residing full-time in Ireland. Respondents were asked before the survey began to state which country they were born in and whether they live in Ireland full-time. Seven interviewers and one study coordinator collected the data. Respondents were recruited into the study through a Facebook study page, emailing the study coordinator, or being contacted through friends and family using snowball sampling. Respondents provided written consent to participate in the study.

The extended survey design and sampling strategy can be seen in more detail in the electronic supplementary material.

2.5 The Nonparametric Modelling Approach

Health state utility valuation data are known to have a complex data structure that can be truncated, skewed and suffer from heteroscedasticity [33]. A nonparametric method allows the utility function to take any form, employing Bayesian hierarchical modelling. This approach can represent complex data, such as health state utility valuation, well and can compare multiple subgroups with small sample sizes as in this study, hence why this method was used for this analysis [34]. We used a similar Bayesian statistical nonparametric model as Kharroubi et al. [35]. Kharroubi et al. [36] have documented how this approach provides a more flexible inference for preference functions when compared with parametric methods.

The modelling aims to estimate the differences between the Polish migrant group and the native Irish group by estimating a general model that explains health state utility valuations in terms of the EQ-5D-5L health descriptive system, respondent background characteristics and migrant status. For $j = 1,2,m$, let $n_j$ be the number of health states that were valued by the jth respondent. An EQ-5D-5L health state is a vector of five numbers, which is represented by the symbol $x$. For $i = 1,2,\ldots,n_j$ and $j = 1,2,\ldots,m$, let $x_{ij}$ be the $i$th health state valued by respondent $j$ and let $y_{ij}$ be the cTTO utility valuation given by respondent $j$ for that health state. Finally, $h(z_j)$ is the vector of covariates for respondent $j$ that includes Polish migrant status ($1 =$ Polish migrant; $0 =$ native Irish), sex ($1 =$ male; $0 =$ female), and age.

The $i$th valuation by respondent $j$ is modelled as shown in Eq. 1:

$$y_{ij} = 1 - \exp(y' h(z_j) + \alpha_j \{1 - u_c(x_{ij})\} + \epsilon_{ij},
$$

where $h(z_j)$ is a vector of functions of the covariates $z_j$ and $\gamma$ is a vector of parameters that, combined, determine the effect of the covariates on the respondents’ valuations; $\alpha_j$ is a zero-mean random respondent effect; $\epsilon_{ij}$ is a zero-mean random error; $u_c(x)$ is a function of the health state vector $x$; and the subscript $c$ represents the respondent’s nationality. The interpretation of $u_c(x)$ is as a base utility function; if the population expectation of $\exp(y' h(z_j) + \alpha_j) \{1 - u_c(x_{ij})\}$ is 1, then $u_c(x)$ is the population utility measure, where $c = 1$ if respondent $j$ is a Polish migrant and $c = 0$ if the respondent is native Irish.

In Eq. 1, every respondent values perfect health as having a utility valuation of 1 (subject to measurement error $\epsilon_{ij}$), no matter what that respondent’s covariates are and regardless of the random effect $\alpha_j$. The covariates and random effects enter model (1) multiplicatively. As a result, they have greater effect the further the base utility is from 1. This implies that we are likely to see more variability in the respondents’ valuations for severe health states than for mild states.
Normal distributions are assigned for the random terms as shown in Eq. 2:

\[ \alpha_j \sim N(y_j^*h(z_j), \tau^2) \text{ and } \varepsilon_{ij} \sim N(\theta, \omega^2), \]  

(2)

where \( \tau^2 \) and \( \omega^2 \) are parameters to be estimated. The distribution of the respondent effect \( y_j = \exp(y_j^*h(z_j) + \alpha_j) \) is then log-normal, representing a skewness that is often observed in health state valuation data. The \( z_j \)'s are centred to ensure that they have zero means, and hence the value of \( z_j \) for a typical person is 1.

The base utility functions for each group are provided below: \( u_0(x) = \text{Irish} \); \( u_1(x) = \text{Polish} \). The relationship between the two base utility functions is modelled as shown in Eqs. 3 and 4.

\[ u_0(x) = \mu_0 + \beta_0^*x + d(x), \]  

(3)

\[ u_1(x) = (\mu_0 + \mu_1) = (\beta_0^* + \beta_1^*)x + d(x), \]  

(4)

The expression \( \mu_0 + \beta_0^*x \) in Eq. 3 states that the underlying utility function \( u_0(x) \) for the Irish respondents will usually behave like a simple linear combination of the elements of the health state description vector \( x \). The coefficients \( \beta_0 \) (which are expected to be negative) represent the rate at which utility generally declines when we increase the severity level in the corresponding dimension of \( x \). The comparable expression in Eq. 4 modifies these underlying trend variables with additional coefficients \( \beta_1 \) to reflect dimension-specific differences between the Polish migrants and native Irish.

The term \( d(x) \) represents a deviation from the simple linear trend that is common to the Polish migrants and native Irish respondents. We treat \( d(x) \) as an unknown function, and therefore in a Bayesian nonparametric framework, it becomes a random variable. It has a zero mean and constant variance and is constrained by a correlation between \( d(x) \) and \( d(x') \) for two different states \( x \) and \( x' \) that decreases as the distance between \( x \) and \( x' \) increases. The effect of this is to assert that if \( x \) and \( x' \) describe very similar health states, their utilities will be approximately the same. This will allow the preference function to vary smoothly as the health state changes. This model allows \( d(x) \) to take any value reflecting a nonparametric specification, and hence the utility functions are not constrained. It is in this regard that we note our model as nonparametric, similar to Kharrouri et al. [35].

It is important to note that the population utility function, is not, in general, the base utility function \( u_c(x) \). The base utility function is the utility function for a person whose respondent effect is 1. This will be the base utility function only if the mean or median respondent effect across the population is 1. Otherwise, the population utility function in country \( c \) is as shown in Eq. 5:

\[ \bar{u}_c(x) = 1 - \bar{a}_c(1 - u_c(x)), \]  

(5)

where \( \bar{a}_c \) is the population mean/median respondent effect for that country and \( u_c(x) \) is the corresponding base utility function. The fact that, in general, \( \bar{a}_c \neq 1 \) and the population utility is not the base utility affects how we interpret the utility slope parameters \( \beta_0 \) for the native Irish and \( \beta_1 \) the Polish migrants. The change in utility for an increase of one level in a particular dimension is represented as an underlying trend, not by the corresponding slope parameter but by the parameter multiplied by \( \bar{a}_c \).

In this paper, we only consider the Bayesian statistical nonparametric model to estimate the differences between both groups, as the validity of the approach has been established in the above-cited papers [35, 36]. It is also worth noting we are not interested in estimating a value set from the responses collected but only to compare health state utility valuations. The models were estimated in MATLAB.

### 3 Results

A total of 242 survey interviews were completed (119 Polish migrants and 123 native Irish). All interviews adhered to the EuroQol protocol [37] and quality control checks ensuring all interviews were deemed to be of high quality with no interviews being dropped from the analysis. Table 1 reports the descriptive statistics of the study and how each variable was described in the analysis.

In Table 2, the observed mean utility values, posterior inference (predicted) mean utility values and standard deviations are reported for both the Polish migrant and native Irish groups for the health states valued. The observed mean cTTO utility valuations ranged from −0.57 (55555) to 0.96 (11211) for the Polish migrant group, and from −0.61 (55555) to 0.96 (11211) for the native Irish group. The predicted mean cTTO utility valuations ranged from −0.42 (55555) to 0.94 (11211) for the Polish migrant group, and from −0.59 (55555) to 0.99 (11211) for the native Irish group.

For the purpose of interpreting the results, the health states are organised by their misery score, which is a proxy measure of the severity of the health state that is calculated by adding together the levels of each health state (for example, health state (33252): 3 + 3 + 2 + 5 + 2 = 15) [38]. For this analysis, health states with a misery score of 10 or less are considered mild, health states with a misery score of > 10 and ≤15 are considered intermediate, and health states with a misery score > 15 are considered severe.

The Polish migrants’ predicted mean utility valuations were lower for mild states compared with the native Irish. For example, for state 11121, the predicted values for the Polish migrant group is 0.9, and 0.92 for the native Irish.
In this study, we examined the association between Polish migrant status and EQ-5D-5L health preferences. Our study is the first to explicitly capture and compare the health preferences of migrants and a host population using a novel application of a smaller design EQ-5D-5L valuation study and targeted sampling strategy. More specifically, this study examined the health preferences of Polish migrants and native Irish, both residing full-time in Ireland. We used the EQ-PVT data collection software and a Bayesian statistical nonparametric model to estimate

### Table 1 Sample descriptive statistics

| Variable                      | Polish migrant [n = 119; 49%] | Native Irish [n = 123; 51%] |
|-------------------------------|-------------------------------|-------------------------------|
| Age, years [continuous]      | 38 (7.6)                      | 34 (15.1)                    |
| Sex [1 = male]                | 22% (0.41)                    | 55% (0.5)                    |
| Married [1 = married]         | 69% (0.46)                    | 27% (0.44)                   |
| Urban household [1 = urban]   | 76% (0.43)                    | 43% (0.5)                    |
| Number of dependents <18 [continuous] | 1.14 (1.04) | 0.35 (0.78) |
| Employed [1 = employed]       | 76% (0.42)                    | 54% (0.5)                    |
| Private health insurance [1 = yes] | 45% (0.5) | 56% (0.5) |
| Third-level education [1 = yes] | 76% (0.43) | 67% (0.47) |
| Self-rated health using EQ-5D-5L misery score [continuous] | 6 (1.86) | 5.6 (1.09) |
| Self-rated health using VAS [continuous] | 84 (11.62) | 85 (10.36) |

Standard deviation is reported in parentheses

VAS visual analogue scale

The covariates and random effects enter the Bayesian model multiplicatively, unlike the traditional regression model which applies additively [35, 36]. As a result, they have more effect the further the base utility is from 1. This implies that we will see greater variability in the respondents’ valuations for poor health states than for good states. This is indeed a feature that is observed in valuation data and is one way in which our model better represents reality than that of conventional classical regression models. We also believe that our modelling is more realistic and allows a clear separation between the overall tendency for Polish respondents to give higher valuations (but with greater differentials for worse health states) from the way in which the Polish population differs from the Irish in its reaction to different health dimensions.

Henry et al. [39] and Shaw et al. [26] both note that a minimally important difference of ≥ 0.1 is a meaningful difference in EQ-5D health state utility valuations. From this analysis, a difference of ≥ 0.1 is found in seven health states between the Polish migrants and native Irish, namely states 31125, 11453, 22245, 14335, 42354, 34444 and 55555. Polish migrants are also predicted to place higher utility valuations on the three most severe health states (42354, 34444 and 55555) where meaningful differences > 0.1 are found.

Figure 1 shows the observed versus predicted mean utility values for the Polish migrant group and Fig. 2 shows the same for the native Irish group. The yellow line in each graph represents the residuals obtained between the observed valuation and the predicted valuation for each group. As can be seen from the respective graphs, the valuations for the Polish migrant group are very similar for mild and intermediate states, with discrepancies emerging for severe states, while the inverse is true for the native Irish valuations. Figure 3 shows the histograms of the conditional posterior distributions of the covariate effects (sex, age and Polish migrant). The histogram of the conditional posterior distribution of the Polish migrant covariate demonstrates it to have an effect as the distribution is centred further away from zero. The result is similar for the age covariate. The effect is weaker for the sex covariate as the distribution is seen to be more closely centred over zero. Both Bayesian statistical nonparametric models achieved a low root mean square error (RMSE) of 0.06 and 0.08 for the Polish migrants and native Irish, respectively. A graph further showing the differences in the predicted utility valuations of both groups can be seen in Fig. 4. The graph shows that differences in predicted utility valuations increases between both groups as the health states increase in severity.

### 4 Discussion

In this study, we examined the association between Polish migrant status and EQ-5D-5L health preferences. Our study is the first to explicitly capture and compare the health preferences of migrants and a host population using a novel application of a smaller design EQ-5D-5L valuation study and targeted sampling strategy. More specifically, this study examined the health preferences of Polish migrants and native Irish, both residing full-time in Ireland. We used the EQ-PVT data collection software and a Bayesian statistical nonparametric model to estimate
the differences in health state utility valuations between the Polish migrants and native Irish.

Preferences for health are noted to differ due to geographic variation [40] and the results from this study show that migrant status can also lead to differences. In this study, the differences between both groups are most apparent in the severe health states, where Polish migrants are predicted to place higher utility values on these states than the native Irish on average. Polish migrants are predicted to place higher utility valuations on 18 of the 30 health states than the native Irish. Of these, Polish migrants are predicted to place meaningfully higher utility valuations of ≥ 0.1 on seven health states, including the three most severe health states. It is justifiable to note that where we find differences of 0.1 or larger in health state utility valuations, these differences can be considered meaningful differences [26, 39]. Of the remaining 12 states, the native Irish are predicted to place higher utility valuations on them when compared with the Polish migrants. However, for each state, the difference is < 0.1 and is therefore not considered meaningful. The largest predicted difference between both groups was in their respective utility valuations of the 55555 state, whereby Polish migrants place a higher utility valuation of 0.17 relative to the native Irish.

The differences in the utility valuations of the health states may be linked to the Polish migrants and native Irish

*Table 2* Inference for all health states by migrant status

| Health state number | Health state profile | Misery score | Polish migrant | Native Irish | Predicted difference (Polish–Irish) |
|---------------------|----------------------|--------------|---------------|--------------|-------------------------------------|
|                     |                      |              | Observed mean | Posterior inference mean | Standard deviation | Observed mean | Posterior inference mean | Standard deviation |                          |
| 1                   | 11111                | 5            | 1             | 1             | 0                     | 1             | 0             | 0                     |                           |
| 2                   | 11112                | 6            | 0.9269        | 0.8941        | 0.037                 | 0.9122        | 0.901         | 0.0329                | −0.007                   |
| 3                   | 11121                | 6            | 0.9137        | 0.9           | 0.0404                | 0.9549        | 0.9197        | 0.0373                | −0.020                   |
| 4                   | 11211                | 6            | 0.9577        | 0.9363        | 0.0384                | 0.9622        | 0.9883        | 0.0339                | −0.052                   |
| 5                   | 12111                | 6            | 0.9062        | 0.8816        | 0.037                 | 0.9463        | 0.9347        | 0.032                 | −0.053                   |
| 6                   | 21111                | 6            | 0.9025        | 0.9023        | 0.0389                | 0.9585        | 0.9557        | 0.0342                | −0.053                   |
| 7                   | 12112                | 7            | 0.8262        | 0.8255        | 0.037                 | 0.85          | 0.8358        | 0.0372                | −0.010                   |
| 8                   | 41231                | 11           | 0.6763        | 0.6461        | 0.0511                | 0.7085        | 0.653         | 0.0511                | −0.007                   |
| 9                   | 31125                | 12           | 0.2587        | 0.2065        | 0.068                 | 0.2354        | 0.0923        | 0.0703                | 0.114                    |
| 10                  | 21514                | 13           | 0.29          | 0.2351        | 0.0644                | 0.2915        | 0.205         | 0.066                 | 0.030                    |
| 11                  | 23323                | 13           | 0.6308        | 0.5773        | 0.0559                | 0.5805        | 0.5264        | 0.056                 | 0.051                    |
| 12                  | 24151                | 13           | −0.0897       | −0.0586       | 0.0779                | −0.0207       | −0.0703       | 0.0743                | 0.012                    |
| 13                  | 35311                | 13           | 0.5662        | 0.5134        | 0.0539                | 0.6646        | 0.592         | 0.0513                | −0.079                   |
| 14                  | 11453                | 14           | 0.0812        | 0.0839        | 0.0757                | 0.1598        | −0.0239       | 0.0802                | 0.108                    |
| 15                  | 13541                | 14           | 0.2037        | 0.1421        | 0.07                  | 0.2024        | 0.1553        | 0.0637                | −0.013                   |
| 16                  | 15224                | 14           | 0.2513        | 0.1464        | 0.0733                | 0.2024        | 0.0775        | 0.0719                | 0.069                    |
| 17                  | 52421                | 14           | 0.4575        | 0.4328        | 0.0563                | 0.5805        | 0.4911        | 0.053                 | −0.058                   |
| 18                  | 22245                | 15           | −0.0167       | 0.0033        | 0.0827                | −0.1171       | −0.1633       | 0.0881                | 0.167                    |
| 19                  | 33252                | 15           | 0.1           | 0.0651        | 0.0705                | 0.1467        | −0.0007       | 0.0737                | 0.066                    |
| 20                  | 51342                | 15           | 0.2256        | 0.1611        | 0.0713                | 0.139         | 0.1268        | 0.0671                | 0.034                    |
| 21                  | 54213                | 15           | 0.1987        | 0.1855        | 0.0663                | 0.2549        | 0.1964        | 0.0623                | −0.011                   |
| 22                  | 14335                | 16           | 0.0075        | −0.0431       | 0.0826                | −0.0671       | −0.1748       | 0.0861                | 0.132                    |
| 23                  | 25432                | 16           | 0.1949        | 0.1387        | 0.0726                | 0.2073        | 0.1335        | 0.0676                | 0.005                    |
| 24                  | 32533                | 16           | 0.4359        | 0.366         | 0.0658                | 0.3768        | 0.3024        | 0.0669                | 0.064                    |
| 25                  | 53134                | 16           | 0.1562        | 0.0841        | 0.0778                | 0.1598        | −0.0112       | 0.0808                | 0.095                    |
| 26                  | 43415                | 17           | −0.0387       | −0.0235       | 0.0766                | 0.0134        | −0.0858       | 0.0754                | 0.062                    |
| 27                  | 44522                | 17           | 0.2913        | 0.2453        | 0.0674                | 0.3268        | 0.2685        | 0.0612                | −0.023                   |
| 28                  | 45143                | 17           | 0.0775        | −0.0669       | 0.0883                | −0.1451       | −0.1409       | 0.083                 | 0.074                    |
| 29                  | 42354                | 18           | −0.1313       | −0.1925       | 0.0926                | −0.3293       | −0.3291       | 0.0934                | 0.137                    |
| 30                  | 34444                | 19           | −0.2167       | −0.1976       | 0.0926                | −0.3463       | −0.3025       | 0.0908                | 0.105                    |
| 31                  | 55555                | 25           | −0.5723       | −0.4235       | 0.1075                | −0.6053       | −0.5927       | 0.113                 | 0.169                    |

Health state 11111 was not valued in the study but was included in the analysis. SD relates to the predicted values.
having different rates of time preference (discount rate). cTTO utility valuations are underpinned by an individual’s rate of time preference, with an individual who places a high utility valuation on health states currently argued to have a lower rate of time preference (i.e. discount future health less heavily) than those who place a low valuation on it [18, 26]. The larger differences in the predicted valuations of the more severe health states compared with the milder states is somewhat to be expected due to the covariates and random effects entering the model multiplicatively, as noted
above. It is worth noting that other factors may influence a migrant’s health preferences, such as their rate of time preference (discount rate), marital status and whether an individual lives in an urban household location [18, 41]. These were not controlled for in this analysis. Kelleher et al. [18] also noted other factors that were found to not be associated with migrant health preferences, which included the length of time a migrant was resident in a host country as well as income levels.

For the first time, this paper presented a novel use of a smaller design EQ-5D-5L valuation study and targeted sampling that can be used to examine the health preferences of minority subgroups. Moreover, the empirical methods of this paper add to the literature as it is the first time the use of a Bayesian statistical nonparametric model was used to estimate health preferences, coupled with a novel application of a smaller design EQ-5D-5L valuation study. For this research, the subgroups under investigation were Polish migrants and native Irish. There is a noted lack of reliable data on migrant health in the literature [42–44]. Collecting migrant health data can pose significant challenges to researchers and healthcare systems for various reasons. For example, being highly mobile groups and hard to locate, migrants can have low response rates to surveys, and data samples uncovered are often highly heterogeneous [45]. These issues may be common to those that arise with other hard-to-reach groups, such as independent young adults. The use of a smaller design EQ-5D-5L valuation study and targeted sampling can help uncover and examine migrant or other subgroups’ health preferences, which can otherwise
be hard to uncover when using the secondary analysis of national valuation studies. The samples collected on the Polish migrants and native Irish emerged unbalance across some sociodemographic characteristics, as described above. More care could have been taken to try to ensure that the data collected on the Polish migrants and native Irish were balanced. Retrospectively, this could have been achieved by collecting data on each group by age group and sex, for example. However, this extra step in the data collection would have made the data collection process more complex.

We argued that health preferences could be a predictor of an individual’s valuation of health, as individuals do not directly value health when they place a utility valuation on a compromised EQ-5D-5L health state that incorporates the time spent in, and quality of, the state. The above-cited papers [14–16] have identified how one values health as an influential factor in determining an individual’s healthcare use and health-related behaviours. By Polish migrants valuing severe health states higher than the native Irish, this may see Polish migrants having different healthcare-seeking practices with different use and need of certain healthcare services. It is conceivable that if Polish migrants value health differently from the Irish resident population, or if any migrant group values health differently from the country’s host population in which they reside, they may engage with healthcare services differently; for example, preventive care uptake. Carrieri and Bilger [14] found that those who place a high value on health are more likely to avail of preventive care than those who do not. The most evident differences in the healthcare use of migrants and host populations is in regard to their respective use of preventive care, with migrants having lower use than host populations [5, 6]. How Polish migrants and native Irish differ in their health preferences and valuation of health may be a factor underpinning some of the disparities in preventive care uptake among migrants and host populations in Europe, as mentioned above.

Kelleher et al. [46] found that health preferences are associated with preventive care uptake. Preventive care uptake and health preferences are both underpinned by an individual’s rate of time preference [18]. Therefore, by the Polish migrants and native Irish having potentially different preferences for health, this may be a factor that can contribute to some of the differences in the preventive care use between migrants and natives. Other factors can also contribute to such differences, such as the formal and informal barriers to healthcare faced by migrants [5].

More research is needed to determine how differences in health preferences, and ultimately differences in the valuation of health, can be linked to differences in healthcare utilisation and health-related behaviours among migrants and host populations. This could be done by collecting health preference data and healthcare utilisation data on a larger sample size of Polish migrants and native Irish or any migrant and native groups. Health preferences could be entered into a model as an independent variable examining its association with a range of healthcare services as used by Kelleher et al. [46]. We believe this paper is an initial building block to examining a potentially critical factor underpinning some of the disparities in healthcare utilisation and health-related behaviours among migrants and host populations.

There are some research limitations to our study that must be noted. Our model did not control for some factors that may determine an individual’s health preferences, such as their rate of time preference, and this is a limitation to the research. We used snowball sampling, therefore a representative sample of Polish migrants or native Irish was not used, given the difficulties associated with trying to obtain a representative sample. When studying migrants, it can be hard to generalise the findings given the heterogeneous nature of migrant samples [45]. The findings from this study cannot be viewed as representative of Polish migrants in Ireland or in general since we did not use a representative sample. Since our survey questions were carried out in English, our sample may be underrepresentative of Polish migrants in Ireland who are not fluent in English. Our sample size was small, however we believe using the Bayesian statistical nonparametric model allowed us to make use of the small number of observations. This particular method is noted to be useful when examining differences in health state utility valuations among subgroups with limited sample sizes [33, 35, 36]. Further research might compare the Bayesian statistical nonparametric model results with parametric methods and extend the number of subgroups under investigation. Further research could also look to compare the health preferences of various other migrant groups and host populations across Europe, especially migrant groups who have recently arrived in a host country.

The results from our study should be viewed as an assessment of the association between Polish migrant status and EQ-5D-5L health preference, and not as representations of value sets for each group more broadly, which was not the intention of this study. Polish and Irish nationals are shown to have different health preferences from their respective national valuation studies [18, 22, 47]. The results from this paper highlight that Polish migrants and native Irish differ in their health preferences for, and valuation of, severe health states. Policymakers charged with the design and implementation of migrant health policy in Ireland and Europe may find the results of this study of relevance. We have shown that migrants can value certain health states differently than host populations, which can provide an understanding of a potential new factor underpinning some of the disparities in healthcare utilisation and health-related behaviours among migrants and host populations in Europe.
5 Conclusion

This study has shown that by using a novel application of a smaller design EQ-5D-5L valuation study coupled with the EQ-PVT data collection software and a Bayesian analytic design, differences in health preferences exist between a sample of Polish migrants and native Irish regarding severe health states. If replicated in other contexts, these findings may contribute to our understanding of migrant health outcomes, health-related behaviours, and healthcare utilisation, as migrants may value health differently than host populations. This study also provides proof of principle for using a smaller design EQ-5D-5L valuation study to explore differences in health preferences among other migrant or minority subgroups, which can otherwise be hard to uncover when using the secondary analysis of national EQ-5D-5L valuation studies.

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Declarations

Conflicts of interest Dan Kelleher, Samer Kharroubi, Edel Doherty, Gianluca Baio, and Ciaran O’Neill have no conflicting interests, monetary or otherwise, that would in any way inhibit the submission and publication of this research. None of the authors have any research or professional conflicts of interest to report.

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Consent to participate All participants consented to participate in this research and signed a consent form.

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Availability of data and material Can be made available upon reasonable request.

Code availability Can be made available upon reasonable request.

Author contributions DK: data curation, data modelling, writing the original draft of the paper and redrafting of the paper. SK: data modelling and redrafting of the paper. ED: redrafting of the paper and supervision. GB: redrafting of the paper. CON: redrafting of the paper and supervision. All authors contributed fairly to this paper.

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