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Research in context

Evidence before this study

Emerging evidence shows that breastfeeding is associated with a reduced maternal long-term risk of hypertension, endometriosis, diabetes, cardiovascular disease, metabolic syndrome, and certain types of cancer (e.g., ovarian and breast cancer), probably by increasing energy expenditure and providing a positive impact on lipid metabolism, glucose homeostasis, and insulin secretion. However, the cumulative role of a woman's lifetime breastfeeding on subsequent risk of mortality, a major public health focus, is poorly known. To fill the data gap, we evaluated the association of lifetime breastfeeding duration with the risk of all-cause and cause-specific mortality among women from two large ongoing prospective cohorts, the Nurses' Health Study (NHS) and the Nurses' Health Study II (NHS II), with continuous follow-up spanning three decades. Because lifestyle factors such as cigarette smoking, nutrition, physical activity, and overweight or obesity are important determinants of mortality and can be influenced by breastfeeding practices, we also investigated the potential interaction between breastfeeding duration and these lifestyle factors across women's reproductive lifespan on mortality risk.

Added value of this study

Among 166,708 parous women from two large prospective cohorts, we found non-linear associations between lifetime breastfeeding duration and the long-term risk of total, cancer, and cardiovascular disease mortality during adulthood. These non-linear associations were further confirmed in the cubic spline, where women who breastfed for a total of 18-24 months experienced the lowest risk. There was no evidence of interactions between breastfeeding duration and pre-pregnancy smoking status, dietary quality, alcohol consumption, physical activity, and overweight or obesity on mortality risk.

Implications of all the available evidence

Our results strengthen and refine the evidence of lifelong benefits of breastfeeding for mothers and highlight the importance for policymakers and social communities to protect, promote, and support breastfeeding and for healthcare professionals to inform its advantages to mothers and infants among the high-risk population with low breastfeeding initiation.

Centres for Disease Control and Prevention. Although these initiatives have been established for more than two decades, the prevalence of exclusive breastfeeding is only about one-third among children younger than 6 months in low- and middle-income countries. In high-income countries, the estimated proportion of breastfeeding at 12 months is lower than 20%.

Breastfeeding offers unrivalled benefits for all children by improving their survival, growth, and lifelong health. Meanwhile, emerging evidence shows that breastfeeding is associated with a reduced maternal long-term risk of hypertension, endometriosis, diabetes, cardiovascular diseases (CVD), metabolic syndromes, and certain types of cancer (e.g., ovarian and breast cancer), probably by increasing energy expenditure and providing a positive impact on lipid metabolism, glucose homeostasis, and insulin secretion.
lifespan. Participants in the NHS II reported their breastfeeding duration in 3 follow-up questionnaires. In 1993, women initially reported their cumulative breastfeeding duration as “cannot remember”, “did not breastfeed”, “<1”, “1–3”, “4–6”, “7–11”, “12–17”, “18–23”, “24–35”, “36–47”, and “≥48 months”. In 1997, a more detailed questionnaire was used to assess the breastfeeding duration of women reporting breastfeeding for each of their first 4 children. Participants who breastfed more than 4 children reported a total of additional months they breastfed for all subsequent children. Women reported additional births after 1997 by completing a similar breastfeeding questionnaire in 2003 when most women had completed their reproductive lifespan.5 To allow comparison with the data from NHS, we calculated the cumulative breastfeeding duration after each birth that the participants reported any breastfeeding for each survey cycle. In 1997 and 2003, women who had breastfed for at least 1 month were also asked “At what month did you start giving formula or purchased milk at least once daily?” and “At what month did you start giving solid food at least once daily?”, which could be reported as “0–2”, “3”, “4–5”, “6–7”, “8–11”, and “≥12 months”. We defined cumulative exclusive breastfeeding duration as the period of time during which women fed infants breast milk without any other liquid or solid food.5 Existing studies have consistently shown that both self-reported breastfeeding initiation and duration are highly reliable.14 Among 146 women whose daughters participated in the NHS, the correlation coefficient of their breastfeeding duration recalled on two occasions 2 years apart was 0.89, indicating high reproducibility.15

Ascertainment of deaths

Deaths in the NHS and NHS II were identified from state vital statistics records and the National Death Index or by reports from next of kin or the postal authorities, which were able to correctly identify >97% of deaths from clinical records.16 Physicians used the Eighth and Ninth Revisions of the International Classification (ICD-8 and ICD-9) to identify the underlying

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Figure 1. Flow diagram for study design, data collection, and exclusion criteria.
cause of death due to CVD and cancer by reviewing autopsy reports, death certificates, and medical records (Appendix Table 2).

Assessment of covariates
In both cohorts, data on race/ethnicity, birthplace, and height were collected at recruitment or during follow-up. Parity (number of pregnancies lasting no less than 6 months), age, weight, age at menopause, postmenopausal hormone use, smoking habit, and history of major chronic conditions were updated biennially since recruitment. We calculated body mass index (BMI) by dividing weight in kilograms by the square of height in meters. Physical activity was assessed every 2–4 years using a validated questionnaire collecting the average time spent per week on walking or hiking, jogging, running, bicycling, swimming, weightlifting, working outdoors, playing squash, racquetball or tennis, engaging in lower-intensity exercise, and performing callisthenics and other aerobic exercises. We estimated the weekly time of doing moderate-to-vigorous activities that required at least 3 metabolic equivalent units per hour. Dietary intake, including alcohol consumption, was collected every 2–4 years using a validated semi-quantitative food frequency questionnaire. We applied the Alternate Healthy Eating Index (AHEI) score to assess the overall quality of the diet. These self-reported lifestyle factors have been validated to be highly reliable in both cohorts.

Statistical analysis
Descriptive analyses were conducted for baseline characteristics according to breastfeeding duration by standardizing to participants’ age distribution. We applied time-dependent Cox proportional hazards regression models to estimate the hazard ratios (HRs) and 95% confidence intervals (CIs) for the associations between lifetime total breastfeeding duration and risk of all-cause and cause-specific mortality in both cohorts, jointly stratified by age in years at the start of follow-up and calendar years of the current questionnaire cycle that were equivalent due to the way we structured the data and formulated the Cox models. Because of the similarity of study design and population, we also pooled the hazard ratios in both cohorts using an inverse variance weighted meta-analysis with the fixed effects model to maximise statistical power. Follow-up times started from the return date of follow-up questionnaire when breastfeeding duration was initially reported until the onset of death or the end of follow-up (June 2016 in NHS and June 2019 in NHS II), whichever occurred first. Our main analysis categorised total breastfeeding duration as ≤3 months (reference), 4–6, 7–11, 12–23, and ≥24 months. Tests for linear trends were evaluated using the Wald test on the continuous breastfeeding duration representing the median values of each category. The potential non-linear relationship between continuous breastfeeding duration and mortality risk was examined using restricted cubic splines with 3 knots since they provided the highest power to detect non-linearity.

Multivariable Cox models were initially adjusted for race/ethnicity (non-Hispanic White or other), age at first birth (<20, 20–24, 25–29, ≥30 years), infertility history (yes or no), woman’s own birth weight (<5.5, 5.5–6.9, 7.0–8.4, 8.5–9.9, ≥10 lbs), oral contraceptive use during puberty (yes or no), woman’s parents worked as professionals, managers, or executives during infancy (yes or no), woman’s parents owned home during infancy (yes or no), maternal history of cancer (yes or no), and parental history of CVD before age 60 years (yes or no). In a secondary model, we additionally adjusted for time-varying pre-pregnancy smoking status (never smoker, former smoker: 1–14, ≥15 cigarettes/d, and current smoker: 1–14, 15–24, ≥25 cigarettes/d), exercise at moderate-to-high intensity (0, 0.01–1.0, 1.1–1.4, 1.5–5.9, or ≥6.0 h/week), alternate healthy eating index (quintiles), alcohol drinking (0, 0.1–4.9, 5.0–14.9, 15.0–19.9, 20.0–29.9, or ≥30 g/d), and body mass index (<21, 21–24.9, 25–29.9, 30–31.9, or ≥32 kg/m²). Time-varying covariates with missing values at a given follow-up cycle were replaced with the information from the most recent cycle (mostly ≤5%); otherwise, we used the missing covariate indicator method to deal with missing values, which has been demonstrated to induce minimal bias in epidemiologic studies.

Stratified analyses were performed to assess the effect modification by time-varying parity (1 or ≥2) and pre-pregnancy BMI, smoking status, dietary quality, alcohol consumption, and physical across women’s reproductive lifespan. Multiplicative interaction was estimated using multiplicative interaction terms between dichotomous breastfeeding duration (≤3 versus >3 months) and these stratified variables based on the Wald test. Several sensitivity analyses were also conducted. First, to assess the potential confounding by parity, we reanalysed the association of lifetime total breastfeeding duration with risk of all-cause mortality by a) additionally adjusting for time-varying parity; and b) using the average lifetime breastfeeding duration per parity by categorising women as <1 (reference), 1–3, 4–6, and ≥7 months. Second, we additionally adjusted for the tier of birth (North, Middle, South, outside of US) to assess the influence of birthplace. Third, we remodelled the Cox proportional hazards regression models by stratifying the analyses jointly by calendar years and participants’ own age in months or broader age groups (<35, 35–39, 40–44, 45–49, 50–54, 55–59, and ≥60 years). Finally, we assessed the influence of pregnancy complications (i.e., gestational diabetes and hypertensive disorders) and exclusive breastfeeding duration among NHS II women who had completed data. All data were analysed using SAS 9.4 for UNIX (SAS Institute Inc).
### Characteristics

| The Nurses’ Health Study (1986) | Lifetime total breastfeeding duration (months) |
|--------------------------------|-----------------------------------------------|
|                                | ≤3                                      | 4–6| 7–11| 12–23| ≥24 |
| Number of women                | 51,168                                   | 10,279| 8333| 10,419| 5765 |
| Age (year), mean (SD)          | 52.8 (7.0)                               | 52.2 (7.2)| 51.3 (7.2)| 51.0 (7.2)| 50.6 (7.2) |
| Non-Hispanic White, %          | 49,936 (97.6)                            | 9988 (97.2)| 8118 (97.4)| 10,146 (97.3)| 5646 (97.8) |
| Woman’s own birth weight less than 2.5 kg, % | 3850 (7.6)                               | 701 (6.8)| 618 (7.3)| 740 (7.0)| 397 (6.8) |
| Oral contraceptive use during puberty, % | 3178 (6.1)                               | 725 (7.1)| 561 (6.9)| 620 (6.2)| 274 (5.1) |
| Parity, mean (SD)              | 3.0 (1.4)                                | 3.1 (1.4)| 3.2 (1.4)| 3.5 (1.4)| 4.5 (1.9) |
| Age at first birth (year), mean (SD) | 25.1 (3.7)                               | 25.1 (3.8)| 25.2 (3.5)| 25.2 (3.5)| 25.2 (3.8) |
| Infertility history, %         | 5162 (10.2)                              | 1098 (10.7)| 872 (10.4)| 958 (8.9)| 509 (8.1) |
| Pre-pregnancy cigarette smoking, % | 12,470 (24.5)                           | 2297 (22.3)| 1643 (19.7)| 1803 (17.4)| 831 (14.3) |
| Pre-pregnancy BMI (kg/m²), mean (SD) | 25.4 (4.9)                               | 25.2 (4.7)| 25.0 (4.5)| 25.2 (4.7)| 25.6 (4.8) |
| Pre-pregnancy diet quality (alternate healthy eating index), mean (SD) | 50.0 (12.1) | 50.0 (12.8)| 50.9 (12.5)| 50.7 (12.6)| 50.1 (12.8) |
| Pre-pregnancy alcohol intake (g/d), mean (SD) | 6.1 (10.6) | 6.5 (10.7)| 6.4 (10.9)| 6.1 (10.4)| 5.2 (9.6) |
| Pre-pregnancy moderate to vigorous intensity exercise (h/wk), mean (SD) | 2.0 (3.1) | 2.3 (3.5)| 2.3 (3.5)| 2.4 (3.5)| 2.4 (3.3) |
| Maternal history of cancer, % | 12,850 (25.1)                            | 2497 (24.3)| 2157 (25.8)| 2600 (24.8)| 1457 (24.9) |
| Parental history of CVD before age 60 years, % | 10,929 (21.4)                           | 2097 (20.4)| 1711 (20.4)| 2099 (20.1)| 1121 (19.2) |
| Woman’s parents worked as professionals, managers, or executives during infancy, % | 14,633 (28.8) | 3236 (31.5)| 2725 (32.4)| 3549 (33.4)| 2030 (34.3) |
| Woman’s parents owned home during infancy, % | 15,565 (30.5) | 3151 (30.6)| 2680 (31.9)| 3385 (32.2)| 1890 (32.6) |

### The Nurses’ Health Study II (1989)

|                                | Number of women | 21,421 | 9011 | 13,232 | 19,305 | 17,775 |
| Age (year), mean (SD)          | 34.7 (5.0)      | 34.2 (4.7)| 34.2 (4.7)| 34.0 (4.6)| 34.0 (4.5) |
| Non-Hispanic White, %          | 19,823 (92.5)  | 8337 (92.5)| 12,377 (93.5)| 18,224 (94.4)| 17,031 (95.8) |
| Woman’s own birth weight less than 2.5 kg, % | 1499 (7.0)      | 588 (6.6)| 849 (6.4)| 1203 (6.2)| 1042 (5.8) |
| Oral contraceptive use during puberty, % | 4592 (21.4)    | 2151 (23.7)| 3085 (23.3)| 4293 (22.2)| 3630 (20.5) |
| Parity, mean (SD)              | 1.5 (1.0)       | 1.6 (1.0)| 1.6 (1.0)| 1.9 (1.0)| 2.3 (1.2) |
| Age at first birth (year), mean (SD) | 24.5 (4.2)     | 25.7 (4.3)| 26.0 (4.2)| 26.0 (3.8)| 25.9 (3.5) |
| Infertility history, %         | 5556 (26.0)    | 2422 (26.9)| 3565 (27.0)| 4679 (24.2)| 3824 (21.4) |
| Pre-pregnancy cigarette smoking, % | 2050 (9.6)    | 808 (8.9)| 964 (7.3)| 1102 (5.7)| 698 (3.9) |
| Pre-pregnancy BMI (kg/m²), mean (SD) | 24.5 (5.3)   | 24.1 (4.9)| 23.9 (4.7)| 23.5 (4.4)| 23.2 (4.1) |
| Pre-pregnancy diet quality (alternate healthy eating index), mean (SD) | 46.2 (10.5) | 46.8 (10.4)| 47.8 (10.5)| 48.0 (10.5)| 48.1 (10.7) |
| Pre-pregnancy alcohol intake (g/d), mean (SD) | 2.9 (5.7) | 3.2 (6.2)| 3.3 (6.2)| 3.0 (5.6)| 2.6 (5.2) |
| Pre-pregnancy moderate to vigorous intensity exercise (h/wk), mean (SD) | 2.9 (5.1) | 2.7 (4.7)| 2.8 (4.7)| 2.8 (4.6)| 2.8 (4.6) |
| Maternal history of cancer, % | 5305 (24.8)    | 2170 (24.3)| 3332 (25.2)| 4849 (25.1)| 4453 (24.9) |
| Parental history of CVD before age 60 years, % | 4815 (22.5)    | 1876 (20.9)| 2751 (20.8)| 3720 (19.3)| 3272 (18.3) |
| Woman’s parents worked as professionals, managers, or executives during infancy, % | 4426 (20.7)    | 2180 (24.2)| 3616 (27.3)| 5689 (29.4)| 5722 (32.2) |
| Woman’s parents owned home during infancy, % | 8504 (39.8)   | 3660 (40.8)| 5681 (42.9)| 8195 (42.4)| 7878 (44.3) |

**Table 1: Age-standardised baseline characteristics of study participants from the Nurses’ Health Study (NHS) and the Nurses’ Health Study II (NHS II) according to lifetime total breastfeeding duration.**

*Values are means (SD) or N (percentages); means (SD) and percentages of all variables except for age are age-standardised.*

*In the NHS, a total of 114 (0.1%), 8419 (9.8%), and 4468 (5.1%) women had missing data on baseline BMI, diet (including alcohol intake), and physical activity, respectively; in the NHS II, a total of 4860 (6.0%), 44,453 (57.9%), 6087 (7.5%), and 996 (1.2%) women had missing data on baseline BMI, diet (including alcohol intake), physical activity, and age at first birth, respectively.*
Role of the funding source
All funding sources had no role in the conduct of the study; collection and analysis of the data; preparation and review of the manuscript; and the decision to submit the manuscript for publication. Y-XW and MA had full access to the data in the study and all authors had full responsibility for the decision to submit for publication.

Results
The baseline characteristics of 166,708 parous women in NHS and NHS II according to breastfeeding duration are shown in Table 1. Together, there were 72,589 (43.5%), 19,290 (11.6%), 21,565 (12.9%), 29,724 (17.8%), and 23,540 (14.1%) women who breastfed for a cumulative total of ≤3, 4−6, 7−11, 12−23, and ≥24 months, respectively. In both cohorts, women who cumulatively breastfed for less than or equal to 3 months had the lowest mean parity and the highest prevalence of smoking.

During 4,705,160 person-years of follow-up, 36,634 deaths were documented in the NHS and NHS II, including 9,880 from cancer and 7,709 from CVD. The pooled crude mortality rate per 1,000 person-years of follow-up for women reporting lifetime total breastfeeding duration of ≤3, 4−6, 7−11, 12−23, and ≥24 months were 16.92, 14.95, 13.41, 12.49, and 11.93, respectively. In the age-adjusted model, a longer lifetime total breastfeeding duration was associated with a lower risk of all-cause mortality in both cohorts (Figure 2). The associations were similar after adjusting for ethnicity, birth weight, age at first birth, infertility history, oral contraceptive use during puberty, parents’ profession and homeownership, maternal history of cancer, and parental history of CVD (model 1) but were attenuated with additional adjustment for pre-pregnancy lifestyle factors (model 2). In the fully adjusted models, the pooled HRs for all-cause mortality during follow-up were 0.95 (95% CI: 0.92 to 0.98), 0.94 (95% CI: 0.91 to 0.98), 0.93 (95% CI: 0.90 to 0.97), and 0.93 (95% CI: 0.89 to 0.97), respectively, for women who breastfed for 4−6, 7−11, 12−23, and ≥24 months, compared to women breastfeeding for less than 4 months over their reproductive lifespan (Figure 2). When cause-specific mortality was explored, a similar pattern of non-linear inverse associations was observed between lifetime total breastfeeding duration and CVD mortality (Table 2). These non-linear inverse associations were

![Figure 2. Hazard ratio (HR) (95% confidence interval (CI)) of all-cause mortality according to lifetime total breastfeeding duration among 166,708 women from the NHS (1986−2016) and NHS II (1989−2019). In age-adjusted Cox proportional hazard regression models, the analyses were stratified jointly by participants’ own age in years at the start of follow-up and calendar years of the current questionnaire cycle. Multivariable Cox model was further adjusted for race/ethnicity (non-Hispanic White; yes or no), woman’s own birth weight (<5.5, 5.5−6.9, 7.0−8.4, 8.5−9.9, or ≥10 lbs), oral contraceptive use during puberty (yes or no), infertility history (yes or no), woman’s parents worked as professionals, managers, or executives during infancy (yes or no), woman’s parents owned home during infancy (yes or no), maternal history of cancer (yes or no), and parental history of CVD before age 60 years (yes or no). Multivariable Cox model 2 was further adjusted for time-varying pre-pregnancy smoking status (never smoker, former smoker: 1−14, ≥15 cigarettes/d, and current smoker: 1−14, 15−24, ≥25 cigarettes/d), alcohol drinking (0, 0.1−4.9, 5.0−14.9, 15.0−19.9, 20.0−29.9, or ≥30 g/dl), exercise at moderate-to-high intensity (0, 0.01−1.0, 1.1−3.4, 3.5−5.9, or ≥6.0 h/week), and alternate healthy eating index (five categories), and body mass index (<21, 21−24.9, 25−29.9, 30−31.9, or ≥32 kg/m2). Tests for linear trends were evaluated using the Wald test on the continuous breastfeeding duration representing the median values of each category. Tests for heterogeneity were conducted using an inverse variance weighted meta-analysis with the random-effects model. NA: not applicable.]
further confirmed in the cubic spline model with adjustment for the same set of covariates (all p-values for non-linearity < 0.01), where women breastfeeding for a total of 16–24 months showed the lowest risk (Figure 3).

Cause-specific analyses were also disaggregated for diagnostic categories with at least 30 deaths attributed to the same diagnostic category in both cohorts (Appendix Table 3). Compared to women who breastfed for ≤3 months showed a lower risk of mortality due to malignant neoplasm of the respiratory system (pooled HR = 0.91; 95% CI: 0.84 to 1.00) and non-malignant diseases of the gastrointestinal system (pooled HR = 0.83; 95% CI: 0.79 to 0.87) and endocrine, nutritional, and metabolic diseases or immunity disorder (pooled HR = 0.77; 95% CI: 0.72 to 0.82), but a slightly higher risk of mortality due to suicide (pooled HR = 1.35; 95% CI: 1.14 to 1.59).

There was no evidence of interaction between lifetime total breastfeeding duration and parity, diet quality, alcohol consumption, and physical activity on the risk of all-cause mortality (Table 3). The pooled HRs for all-cause mortality in both cohorts were substantially unchanged when we used the average duration of total breastfeeding per parity, when we additionally adjusted for the Tier of birth or time-varying parity, and when we stratified the analyses jointly by calendar years and participants’ own age in months or broader age groups (Appendix Table 4–7). Among NHS II women, the non-linear inverse associations between breastfeeding duration and all-cause mortality persisted when we additionally adjusted for the history of gestational diabetes and hypertensive disorders (Appendix Table 8) and when we used exclusive breastfeeding duration (Appendix Table 9).

Discussion
In these two large cohorts involving 166,708 parous women, we found that a longer lifetime total breastfeeding duration was associated with a modestly lower risk of all-cause mortality later in life. This inverse association appeared to be independent of parity and persisted when we used the average duration of breastfeeding per...
Figure 3. Pooled dose-response associations between lifetime total breastfeeding duration and risks of all-cause (a) and cause-specific mortality (b-d) among 166,708 women from the NHS (1986–2016) and NHS II (1989–2019). Cubic spline models were adjusted for participants’ own age in years (continuous), race/ethnicity (non-Hispanic White; yes or no), age at first birth (<20, 20–24, 25–29, ≥30 years), woman’s own birth weight (<5.5, 5.5–6.9, 7.0–8.4, 8.5–9.9, ≥10 lbs), oral contraceptive use during puberty (yes or no), infertility history (yes or no), woman’s parents worked as professionals, managers, or executives during infancy (yes or no), woman’s parents owned home during infancy (yes or no), maternal history of cancer (yes or no), parental history of CVD before age 60 years (yes or no), and time-varying pre-pregnancy smoking status (never smoker, former smoker: 1–14, ≥15 cigarettes/d, and current smoker: 1–14, 15–24, ≥25 cigarettes/d), alcohol drinking (0, 0.1–4.9, 5.0–14.9, 15.0–19.9, 20.0–29.9, or ≥30 g/d), exercise at moderate-to-high intensity (0, 0.01–1.0, 1.1–3.4, 3.5–5.9, or ≥6.0 h/week), and alternate healthy eating index (five categories), and body mass index (<21, 21–24.9, 25–29.9, 30–31.9, or ≥32 kg/m²).
eventually to mortality. Animal studies have
tribute to an elevated risk of CVD and cancer, and
models, where women who breastfed for a total of 18
ear associations were further confirmed in cubic spline
observed for CVD and cancer mortality. These non-lin-
similar pattern of non-linear inverse associations was
child. When cause-specific mortality was explored, a
similar pattern of non-linear inverse associations was
observed for CVD and cancer mortality. These non-lin-
early associations were further confirmed in cubic spline
models, where women who breastfed for a total of 18
24 months experienced the lowest risk.

The observed inverse association between breastfeed-
ing duration and mortality risk is biologically plausible
given the pregnancy-associated physiological changes
such as fat accumulation that are essential for the devel-
oment of the fetus and anticipation of lactation. This
energy storage is characterised by well-described
changes in visceral fat, insulin resistance, insulin pro-
duction, and circulating lipid levels,
which could con-
tribute to an elevated risk of CVD and cancer, and
eventually to mortality. Animal studies have
demonstrated that breastfeeding plays an important role
in mobilizing stored fat,
reestablishing glucose homeostasis,
regulating insulin secretion,
and promoting lipid metabolism after delivery.
Population evidence from observational studies also suggests that
breastfeeding is associated with improved glucose metab-
olism and pancreatic beta-cell function,
more favourable lipid profiles,
and reduced weight loss and risk of
metabolic diseases.

Meanwhile, several randomised
controlled trials have demonstrated that greater intensity
of lactation is associated with greater postpartum weight
loss.
Collectively, these findings suggest that breastfeed-
ing plays a critical role in “resetting” maternal metab-
olism by increasing metabolic expenditure, which reduces
the maternal risk of developing CVD and cancer and
eventually the risk of mortality later in life.

| Stratified factors across the reproductive lifespan | Lifetime total breastfeeding duration |
|---------------------------------------------------|--------------------------------------|
|                                                   | ≤3 months | >3 months |
| AHEI diet quality score                           |           |           |
| Top 40% (n=13,821 deaths)                         | 1.00 [Reference] | 0.95 (0.92, 0.99) |
| Bottom 60% (n=22,813 deaths)                     | 1.00 [Reference] | 0.93 (0.91, 0.96) |
| P for multiplicative interaction<sup>b</sup>      | 0.32      |           |
| Smoking status                                    |           |           |
| Never (n=14,074 deaths)                           | 1.00 [Reference] | 0.95 (0.91, 0.98) |
| Current/ever (n=22,560 deaths)                   | 1.00 [Reference] | 0.92 (0.90, 0.95) |
| P for multiplicative interaction<sup>b</sup>      | 0.17      |           |
| BMI                                               |           |           |
| <25 kg/m<sup>2</sup> (n=16,187 deaths)           | 1.00 [Reference] | 0.97 (0.94, 1.00) |
| ≥25 kg/m<sup>2</sup> (n=20,391 deaths)           | 1.00 [Reference] | 0.92 (0.89, 0.94) |
| P for multiplicative interaction<sup>b</sup>      | 0.18      |           |
| Physical activity                                 |           |           |
| ≥30 min/day (n=4387 deaths)                      | 1.00 [Reference] | 0.94 (0.88, 1.00) |
| <30 min/day (n=32,247 deaths)                    | 1.00 [Reference] | 0.93 (0.91, 0.95) |
| P for multiplicative interaction<sup>b</sup>      | 0.99      |           |
| Alcohol consumption                               |           |           |
| 5-15 g/day (n=6013 deaths)                       | 1.00 [Reference] | 0.94 (0.89, 0.99) |
| <5 or ≥15 g/day (n=30,621 deaths)                | 1.00 [Reference] | 0.94 (0.92, 0.96) |
| P for multiplicative interaction<sup>b</sup>      | 0.92      |           |
| Parity                                            |           |           |
| 1 (n=3395 deaths)                                 | 1.00 [Reference] | 0.88 (0.81, 0.96) |
| ≥2 (n=33,239 deaths)                             | 1.00 [Reference] | 0.95 (0.93, 0.98) |
| P for multiplicative interaction<sup>b</sup>      | 0.057     |           |

<sup>a</sup> Cox proportional hazards regression models were stratified jointly by participants’ own age in years at the start of follow-up and calendar years of the current questionnaire cycle, with adjustment for race/ethnicity (non-Hispanic White; yes or no), age at first birth (<20, 20–24, 25–29, ≥30 years), woman’s own birth weight (<5.5, 5.5–6.9, 7.0–8.4, 8.5–9.9, ≥10 lbs), oral contraceptive use during puberty (yes or no), infertility history (yes or no), woman’s parents worked as professionals, managers, or executives during infancy (yes or no), woman’s parents owned home during infancy (yes or no), maternal history of cancer (yes or no), parental history of CVD before age 60 years (yes or no), and time-varying pre-pregnancy smoking status (never smoker, former smoker: 1–14, ≥15 cigarettes/d, and current smoker: 1–14, ≥15–24, ≥25 cigarettes/d), alcohol drinking (0, 0.1–4.9, 5.0–14.9, 15.0–19.9, 20.0–29.9, or ≥30 g/d), exercise at moderate-to-high intensity (0, 0.1–1.9, 2.0–3.4, 3.5–5.9, or ≥6.0 h/week), and alternate healthy eating index (five categories), and body mass index (<21, 21–23, 23–24, 24–25, 25–30, ≥31 kg/m<sup>2</sup>), excluding the stratified variables.

<sup>b</sup> Multiplicative interaction was estimated using multiplicative interaction terms between dichotomous breastfeeding duration and these stratified variables based on the Wald test.
To date, very few studies have examined the association between breastfeeding duration and subsequent risk of mortality, although breastfeeding duration has been associated with a reduced risk of various maternal health outcomes, including hypertension, diabetes, CVD, metabolic syndromes, and certain types of cancer (e.g., ovarian and breast cancer). In support of our findings, Nguyen and colleagues reported that ever breastfeeding was associated with a lower risk of incident CVD hospitalization and mortality among 100,864 parous Australian women aged ≥45 years. Similarly, in a Norwegian prospective cohort study consisting of 21,889 parous women aged 30–85 years, Fagerhaug and colleagues reported a higher risk of CVD mortality among women younger than 65 years who never breastfed. A more recent study consisting of >320,000 European women aged 25–70 years at entry, also found a lower risk of all-cause and circulatory disease mortality among women who ever versus never breastfed. In contrast to our findings, they did not find any convincing association between breastfeeding duration and cancer mortality.

Our results were also contrasted with the results from a Chinese cohort showing that breastfeeding duration was unrelated to CVD mortality risk among 267,400 female textile workers aged ≥30 years at enrollment. The differences between studies could partly be explained by the differences in population characteristics, study design, and sample size. For instance, most of these previous studies determined breastfeeding duration at a single time point by recalling at recruitment when a large portion of women hadn’t completed their reproductive careers, which may have resulted in exposure misclassification. Besides, previous studies mostly classified participants into a few exposure categories (e.g., ever vs. never), which may have been insufficient to detect potential non-linear relationships. Finally, these previous studies lacked detailed information on various relevant confounders such as pre-pregnancy lifestyle factors that are important determinants of mortality and can be influenced by breastfeeding practices.

In our present study, we did not observe any evidence of interaction between breastfeeding duration and cigarette smoking, nutrition, physical activity, and overweight or obesity on mortality risk. However, the associations between breastfeeding duration and total and cause-specific mortality were attenuated when we additionally adjusted for these pre-pregnancy lifestyle factors. We found non-linear associations between lifetime breastfeeding duration and the risk of total, cancer, and CVD mortality. A recent meta-analysis consisting of 18 cohort studies reported a U-shaped association between parity and all-cause mortality where women with three to four live births had the lowest risk, suggesting that high parity may have offset the protective effect of breastfeeding. Interestingly, the inverse association between lifetime breastfeeding duration and mortality persisted when we restricted the analysis among women reporting single parity and when we used the average duration of breastfeeding per child. We also observed a slightly higher risk of suicide mortality among women who cumulatively breastfed for >3 months than those breastfeeding for ≤3 months. However, the possibility of a chance finding cannot be fully ruled out given the limited number of suicide deaths (n=143). More studies are needed to verify our findings and investigate the mechanisms underlying these observed associations.

The strengths of our current study include its prospective design, extensive follow-up period, large population size, sufficient and valid death cases, high response rate, and repeated collection of various reproductive and lifestyle factors across most of the women’s reproductive lifespan. Our study also has some limitations. First, we relied on self-reported breastfeeding duration to assign participants’ exposure, which may have resulted in exposure misclassification, although validation studies have consistently shown high validity of self-reported breastfeeding duration. In this case, however, such misclassification would be non-differential with respect to deaths because of our prospective study design, which would have biased risk estimations toward the null. Second, our participants were female nurses who had homogeneous race and ethnicity, educational attainment, and socioeconomic status, which may limit the generalizability of our findings, particularly for women in low-middle income countries. However, there is no evidence showing that the influence of breastfeeding on maternal health differs by race, ethnicity, or socioeconomic status. Third, our analyses used observational data, which cannot demonstrate any causal effects. However, as it is ethically problematic to randomise parous women to different categories of breastfeeding duration, a randomised controlled trial is infeasible and analyses of well-designed cohort studies such as ours are needed to advance our understanding of the lifelong benefits of breastfeeding for parous mothers.

Over the past two decades, the recommendations established by WHO and other health organizations have promoted the increasing trends in breastfeeding rates on a global scale. However, the prevalence of exclusive breastfeeding remains far below international targets of 50% by 2025 and 70% by 2030, particularly in low-income and middle-income countries. In the present study involving 166,708 parous women from two large prospective cohorts, we found that a longer lifetime breastfeeding duration was associated with a lower risk of mortality in later life in a non-linear manner, where women who breastfed for a total of 18–24 months experienced the lowest risk. Our results strengthen and refine the evidence of lifelong benefits of breastfeeding for parous mothers and highlight the importance for policymakers and social communities to
protect, promote, and support breastfeeding and for healthcare professionals to inform its advantages to mothers and infants among the high-risk population with low breastfeeding initiation.

Contributors
Y-XW analysed and drafted the manuscript. Y-XW and JEC were involved in the study conception and design. MA verified the underlying data and conducted a technical review. JEM and JEC obtained funding for the study. Y-XW, MA, JW-R, EJM, LW, SAM, and JEC participated in the interpretation of the results and critical revision of the manuscript. Y-XW, JEC, and MA had full access to all the data in the study. All authors accept the responsibility to submit for publication.

Data sharing statement
Data described in the manuscript, code book, and analytic code will not be made publicly available. Further information including the procedures for obtaining and accessing data from the Nurses’ Health Studies II is described at https://www.nurseshealthstudy.org/researchers (email: nhsaccess@channing.harvard.edu).

Declaration of interests
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Supplementary materials
Supplementary material associated with this article can be found in the online version at doi:10.1016/j.eclinm.2022.101691.

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