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Neurodevelopmental Outcomes among Extremely Premature Infants with Linear Growth Restriction

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Abstract

Objective—To compare neurodevelopmental outcomes in linear growth-restricted (LGR) infants born <29 weeks with and without weight gain out of proportion to linear growth

Study Design—We compared 2-year neurodevelopmental outcomes between infants with and without LGR and between LGR infants with and without weight gain out of proportion to linear growth. The outcomes were Bayley-III cognitive, motor, and language scores, cerebral palsy, Gross Motor Function Classification System (GMFCS) level ≥2, and neurodevelopmental impairment.

Result—1227 infants were analyzed. LGR infants were smaller and less mature at birth, had higher BMI, and had lower Bayley-III language scores (82.3 vs 85.0, p<0.05). Among infants with LGR, infants with high BMI had lower language scores compared to those with low-to-normal BMI (80.8 vs 83.3, p<0.05), and were more likely to have GMFCS level ≥2 and neurodevelopmental impairment.

Conclusion—Among infants with LGR, weight gain out of proportion to linear growth was associated with poorer neurodevelopmental outcomes.
Introduction

The association of postnatal growth with neurodevelopmental outcomes in preterm infants is well established. Greater gains in weight, length, head circumference (HC), and body-mass index (BMI) during infancy have all been associated with improved neurodevelopmental scores measured during childhood. Greater attention has been paid to promoting weight gain and head growth in premature infants, although some data exist suggesting that linear growth is also important when considering long-term outcomes. For example, in very low birthweight infants, greater length z-score at discharge correlated with higher Bayley-III language score at 24 months, and more rapid linear growth between term and 4 months’ corrected age was associated with lower odds of IQ < 85 at 18 months of age in preterm infants.

In preterm infants, postnatal linear growth restriction (LGR), represented by a decline in length z-score from birth to discharge, is common. Consequently, it is not uncommon for these infants to gain weight out of proportion to their gain in length, with a resulting increase in BMI. Although an increase in BMI between one week of age and term-equivalent age has been associated with improved neurodevelopmental outcomes in moderately preterm infants, it is unclear whether weight gain out of proportion to linear growth (i.e. high BMI), particularly in the setting of linear growth restriction, confers benefit for extremely preterm infants. The importance of following linear growth is that this measurement is a surrogate for lean body mass accrual and thus represents organ growth and differentiation. Thus, linear growth, rather than weight gain, may better represent brain growth and development in premature infants. The objective of this study was to compare neurodevelopmental outcomes among extremely preterm infants with and without LGR (measured from birth to discharge/120 days of age) and with high versus low-to-normal BMI (determined at discharge/120 days of age) in a population experiencing LGR. We hypothesized that, among infants with LGR, weight gain out of proportion to linear growth (i.e. high BMI) would not be associated with improved neurodevelopmental outcomes assessed at 2 years of age.

Methods

This is an observational, retrospective study of prospectively collected data as part of the National Institute for Child Health and Human Development’s Neonatal Research Network Generic Database (GDB) and Follow-Up studies. Infants 23 0/7 to 28 6/7 weeks gestation or with a birthweight of 401 to 1000 grams born 7/1/2012 to 6/30/2014 at participating NICHD Neonatal Research Network sites and who had neurodevelopmental assessment at 22–26 months’ corrected age (CA) were considered for inclusion. Infants were excluded if they carried a diagnosis affecting linear growth (ex. skeletal dysplasia), had length z-score < −2 (which may reflect an unknown inherited condition affecting linear growth) or HC z-score > 2 at birth, or had missing growth or outcome data. All data collection for the birth hospitalization continued until infants reached “status”, defined as hospital discharge or 120 days, whichever came first.

Length, weight and head circumference were measured according to the local practice in each neonatal intensive care unit (NICU); there is no standard practice or required method.
for measuring length among NRN sites. Although the pattern of linear growth restriction has been previously described, there is no well-accepted definition. Standard deviation scores, or z-scores, have been used to describe patterns of growth in extremely preterm infants, and may be better suited to do so when compared with traditional growth percentiles. Therefore, we chose to define LGR as being present when the length z-score at status was more than one point below the length z-score at birth, i.e.

\[
LGR = \text{Length z-score at status} - \text{length z-score at birth} < -1
\]

The Fenton growth chart served as the reference for growth percentiles and corresponding z-scores, which were determined using the bulk calculator provided at http://www.ucalgary.ca/fenton/2013chart. BMI at status was used to determine high ( ≥75th percentile) versus low-to-normal ( < 75th percentile) weight gain relative to linear growth. BMI percentile was determined using the Olsen (CA ≤ 41 weeks) and WHO (CA > 41 weeks) BMI curves. Data collected included maternal/demographic characteristics, time to first feed and full feeds (120 mL/kg/day), duration of parenteral nutrition and postnatal steroid use. Additional data for in-hospital morbidities included rates of culture-positive sepsis (both early and late), necrotizing enterocolitis requiring surgery, grade 3–4 intraventricular hemorrhage (IVH), periventricular leukomalacia (PVL), severe bronchopulmonary dysplasia (BPD, defined as receiving supplemental O2 or positive-pressure ventilation for the first 28 days of life and receiving supplemental O2 ≥ 0.30 or positive-pressure ventilation at 36 weeks PMA or discharge if discharged before 36 weeks), retinopathy of prematurity (ROP) requiring surgery, and patent ductus arteriosus (PDA) requiring surgery. Standardized neurodevelopmental examinations were performed at 22–26 months CA by certified examiners at each NRN center. Gross motor function was assessed with the Gross Motor Function Classification System (GMFCS) in all children. CP was defined as abnormal tone or reflexes in at least 1 extremity and abnormal control of movement or posture to a degree that interferes with age-appropriate activity. Children with CP were defined as having moderate-to-severe CP if they had a GMFCS level ≥ 2. The Bayley Scales of Infant Development, Third Edition were performed by trained, certified examiners.

The primary outcomes for this study were Bayley-III composite cognitive, motor, and language scores in LGR infants with high and low-normal BMI. Additional outcomes compared included the incidence of moderate/severe cerebral palsy, gross motor functional classification scale (GMFCS) level ≥ 2 and neurodevelopmental impairment (NDI), defined as Bayley-III composite cognitive score < 70, Bayley-III composite motor score < 70, GMFCS level ≥ 2, bilateral blindness (< 20–200), or hearing impairment (permanent hearing loss that does not permit the child to understand directions of the examiner and communicate ± amplification with cochlear implants or hearing aids). These same outcomes were also compared between infants with and without LGR.

Bivariate analyses were used to compare demographic and clinical characteristics between LGR infants with excessive versus low-normal BMI, as well as between LGR and non-LGR infants using t-tests or Mann Whitney-U for continuous variables and exact tests for categorical variables. Generalized linear models and logistic regression were used to assess
the association of neurodevelopmental outcomes with LGR and with BMI among infants with LGR. Models were adjusted for maternal socioeconomic status (using maternal insurance status), infant gender, gestational age at birth, birth weight, small-for-gestational age, and center (model 1). A secondary analysis (model 2) also corrected for postnatal head growth, assessed as postnatal head-sparing, which has been previously defined as: HC z-score at status – HC z-score at birth ≥ −1.9 All analyses were performed using SAS version 9.4 with a p-value < 0.05 indicating statistical significance.

Results

Of the 2257 infants born 7/1/12 through 6/30/14, 595 (26%) died and an additional 435 infants were excluded (Supplemental Figure). The most common reason for exclusion was missing growth or outcome data. From the remaining 1227 infants, 912 (74%) met the definition for LGR. For infants with LGR, 353 (39%) had a BMI ≥ 75th percentile at status.

Maternal and infant characteristics are compared between infants with and without LGR, as well as among LGR infants with high versus low-normal BMI as shown in Table 1. Maternal characteristics were similar between infants with and without LGR, as well as among LGR infants with high and low-normal BMI. Compared with infants without linear growth restriction, LGR infants had lower GA, BW, HC and BMI, but higher length, at birth and received more days of parenteral nutrition. Among infants with LGR, those with BMI at status ≥75th percentile were younger, had a higher BW and BMI at birth, and were less likely to be small for gestational age when compared with those with BMI < 75th percentile. LGR infants with a high BMI at status reached full feeds sooner and received fewer days of parenteral nutrition compared with LGR infants with a low-normal BMI.

Infant anthropometrics at status are shown in Table 2. Infants with and without LGR were of similar age at status, but LGR infants had lower weight, length and head circumference, and higher BMI compared to infants without LGR. Among infants with restricted linear growth, those with a high BMI at status were older, had a larger HC, and, as expected, had a significantly greater weight at status; the difference in mean weight at status was nearly 600 grams.

The distributions of length (for all infants) and BMI curves (for LGR infants only) at birth and status are shown as density curves with histograms in Figure 1. As expected for infants with LGR, the distribution of length percentiles shifted leftward from birth to status, representing a negative change in z-score from birth to status, while the distribution in length percentiles for infants without LGR was preserved (Figure 1A). The distribution of BMI at birth for LGR infants with high versus low-normal BMI was relatively similar, and at status the difference noted in these distributions, shown in Figure 1B, represented weight gain out of proportion to linear growth (i.e. high vs low-normal BMI).

Results of bivariate analyses comparing in-hospital morbidities between groups are shown in Table 2. Compared to infants without LGR, those with LGR had higher rates of several morbidities including sepsis, grade 3–4 IVH and ROP requiring surgery. Among infants with
linear growth restriction, those with an elevated BMI at status were more likely to be diagnosed with BPD and receive postnatal steroids than those infants with low-normal BMI. Bayley-III composite scores at 22–26 months are shown in Figure 2 for all groups. In unadjusted analyses, mean Bayley-III motor and language scores were significantly lower in infants with LGR versus those without LGR (86.3 vs 88.7 and 82.3 vs 85.0, respectively, p<0.05 for both comparisons). Among infants with LGR, mean Bayley-III language score was significantly lower in infants with a high BMI at status compared to those with a low-normal BMI (80.8 vs 83.3, p=0.03). In the initial multivariable analysis, only the difference in Bayley-III language score between infants with and without LGR remained significantly different (adjusted mean difference −2.35, 95% CI −4.49 to −0.21, p=0.03) (Figure 2). Results of secondary outcomes are shown in Table 3. As noted, there were no significant differences in these outcomes between infants with and without LGR, and for LGR infants with high versus low-normal BMI in the initial multivariate analysis. However, after controlling for postnatal head growth, LGR infants with high BMI had lower observed Bayley-III language scores (adjusted mean difference −2.33, 95% CI −4.56 to −0.10, p=0.04) and higher odds of GMFCS level ≥ 2 (OR 1.69, 95% CI 1.05 to 2.72, p=0.03) and NDI (OR 1.50, 95% CI 1.02 to 2.20, p=0.04) compared with LGR infants with low-normal BMI. After addition of postnatal head growth to the multivariate model the difference in Bayley-III language scores between infants with and without LGR was no longer statistically significant (adjusted mean difference −2.07, 95% CI −4.25 to 0.11, p=0.06).

Discussion

In this study of premature infants born at <29 weeks’ gestation, we have shown that acquired LGR between birth and hospital discharge is common, present in three quarters of infants in the data set. LGR is associated with poorer language outcomes at 2 years’ corrected age. Among infants with LGR, additional weight gain alone does not ameliorate the neurodevelopmental deficits associated with linear growth restriction, and in fact may confer a disadvantage.

While the association between postnatal growth failure and poorer neurodevelopmental outcome in preterm infants has been well-established, growth failure has been largely described using weight gain and sometimes head circumference. As a result, greater attention is often paid to achieving weight gain, and perhaps head growth, with relatively little emphasis placed on achieving or maintaining linear growth. For example, the New York State Perinatal Quality Collaborative, working with 18 regional perinatal centers, has focused on reducing extrauterine growth restriction among infants < 31 weeks gestation since 2010. However, the collaborative focused on reducing the percent of infants discharged with weight < 10th percentile and improving head growth (measured as the difference in HC z-scores between birth and discharge), with no stated goal to improve linear growth.

In a cohort of 62 very low birthweight infants, Ramel et al. reported a positive correlation between length z-score at NICU discharge and Bayley-III language score at 24 months’ CA, as well as a correlation between greater length z-score at 4 and 12 months’ CA and

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improved cognitive scores at 24 months’ CA after controlling for weight and head circumference z-scores. Similarly, more rapid linear growth between term and 4 months’ CA was associated with improved neurodevelopmental scores at 18 months in infants born < 33 weeks. Although these studies suggest that improved linear growth both before and after term-equivalent age is linked with neurodevelopment, to our knowledge ours is the first study to report on linear growth during hospitalization and neurodevelopmental outcomes in a large cohort comprised exclusively of extremely low gestational age newborns.

Whether greater weight gain relative to linear growth, reflected by an increase in BMI, confers neurodevelopmental benefit is still not clear. In infants born < 33 weeks, an increase in BMI from birth to term has been associated with better neurodevelopmental outcomes at 18 months’ CA, and BMI gains later in infancy have been associated with lower odds of IQ < 85 later into childhood. Moreover, the study by Belfort et al noted that the association between increasing BMI and improved neurodevelopmental outcomes was isolated to infants with birth weight < 1250 grams. In our study, LGR infants who experienced greater weight gain relative to linear growth, and thus an increase in BMI, did not experience improved neurodevelopmental outcomes at 2 years of age when compared with LGR infants with weight gain more proportional to linear growth (low-normal BMI). The exact reason for these conflicting results regarding BMI and neurodevelopment is not entirely clear, but may be due to the fact that our study focused specifically on outcomes in LGR infants. An increase in BMI, by definition, indicates a greater increase in weight relative to length. Anthropometric curves suggest an increase in BMI from birth to term in preterm infants is to be expected. The pattern of BMI increase associated with optimal outcomes however, such as weight gain with appropriate linear growth versus weight gain in the setting of linear growth restriction, is not well-established. In fact, weight gain out of proportion to linear growth may be undesirable for long-term health. On multivariable analysis, LGR infants with high BMI at status had lower Bayley-III language scores and were more likely to have GMFCS ≥ 2 and NDI when compared with LGR infants with low-normal BMI.

Additionally, greater gains in BMI both before and after term have been associated with greater risk of obesity at 8 and 18 years of age in preterm infants. Identifying and promoting patterns of growth that optimize outcomes in preterm infants, particularly those born extremely premature, is critical to provide the highest level of care for these infants.

While the exact reasons for poor linear growth, as well as different patterns of weight gain in the setting of poor linear growth, remain unknown, it is reasonable to consider potential mechanisms. Certainly, nutritional mediators impact infant growth, including gains in length. Both longer time to reach full enteral feeds and greater caloric deficit during hospitalization have been associated with lower length z-score at discharge in VLBW infants. Nutritional interventions that promote early enteral feeding and target macronutrient recommendations have been successful in improving linear growth in preterm infants. Linear growth, a surrogate for lean body mass accrual, and thus organ growth, may better represent brain growth and development in premature infants. Specific nutrient and energy deficits can impact the growth and development of different brain regions, resulting in altered brain structure and function. More specifically, these nutrient and energy deficits, which occur commonly in premature infants, may result in regionalized effects on brain development. We speculate that critical windows of development likely exist
(before term, early infancy, and later infancy/childhood), which may help explain why certain growth parameters, such as linear growth, do not consistently correlate with neurodevelopmental outcomes across all previously investigated “windows”. Our database (GDB) does not have comprehensive nutritional information for these infants from which to draw conclusions about the role of nutrition in our findings.

In our population, LGR infants with elevated BMI at status were more likely to have BPD compared with LGR infants with low-normal BMI, and all LGR infants had a higher rate of culture-positive sepsis compared with infants without LGR. Thus, the role inflammation may play in poor linear growth should be considered. The accrual of lean body mass, measured indirectly by linear growth, is altered by systemic inflammation. A sepsis model in rats, mediated by TNF-α, results in decreased protein synthesis in skeletal muscle.\textsuperscript{20, 21} Better outcomes in anthropometrics and bone growth have been observed in preterm infants with lower levels of inflammatory mediators.\textsuperscript{22, 23} While the incidence of BPD was significantly higher in LGR infants with high BMI, due to this hypothesis that the same underlying mechanism (i.e. systemic inflammation) may contribute to both LGR with high BMI and BPD, we have chosen not to additionally adjust models for BPD, as doing so would detract from the effect of interest. Moreover, though BPD has been associated with poorer neurodevelopmental outcomes in extremely preterm infants, conflicting data on the true influence of BPD on neurodevelopment exist, with some studies failing to demonstrate BPD as an independent influencer of neurodevelopmental impairment.\textsuperscript{25, 26, 27, 28}

This study has several limitations. The definitions of both LGR and low-normal versus elevated BMI, although established \textit{a priori}, were arbitrary. The definition of LGR does, however, parallel a similar cut-off defining postnatal head sparing (change in head circumference z-score $\geq -1$ from birth to status) among growth-restricted premature infants that we have found to be associated with improved neurodevelopmental outcome.\textsuperscript{9} Length itself is inherently difficult to measure, and was not consistently measured using a length board, considered the most accurate method.\textsuperscript{24} Preliminary evaluation from the University of Rochester NICU has shown poor agreement between length measured serially by tape measure and length board, with only 62\% of 97 paired measurements agreeing to within 1 cm (unpublished data). However, the same evaluation showed no skew of tape measure measurements toward higher or lower values relative to length board determinations. Thus, tape measure lengths, while imprecise, were accurate (i.e., showed no systematic bias). As a result, while there may have been some misclassification of infants as having or lacking LGR at the borders of the current data, we would not expect any systematic misclassification. Another potential limitation exists with the choice of growth curves used to measure weight, length, HC and BMI. Slight differences do exist in the predicted percentiles for weight and length between the Fenton and Olsen curves, which may have introduced systematic bias. The Fenton curves, rather than Olsen, were chosen to calculate weight, length and HC z-scores because they were derived from a more comprehensive and diverse patient population and include infants born as young as 22 weeks gestation.\textsuperscript{10} However, the Fenton anthropometric data do not include a weight-for-length comparison (such as BMI), and thus the Olsen BMI curves were used to calculate BMI percentiles for those infants $\leq 41$ weeks CA. In this study, we expect any systematic bias to be slight, and thus believe the benefits of using the chosen growth curves outweigh this risk. Poor linear
growth in our population was associated with poor HC growth, which we and others have shown to predict poorer neurodevelopmental outcome. However, secondary analyses accounting for HC growth showed little difference in outcomes, suggesting that linear growth is an independent risk factor for poor neurodevelopmental outcome. Finally, the retrospective nature of the dataset and the limited nutritional data restrict both the ability to draw causal conclusions and the possibility of identifying potential mediators of LGR. On the other hand, conclusions drawn from this large dataset representing multiple centers is likely to be generalizable to other groups of similar infants.

**Conclusion**

Among extremely preterm infants with LGR, weight gain out of proportion to linear growth did not result in improved neurodevelopmental outcomes, and was associated with poorer performance on some measures. Improved linear growth during hospitalization is associated with improved language development; this finding deserves further exploration.

**Supplementary Material**

Refer to Web version on PubMed Central for supplementary material.

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Figure 1.
Graphical depiction of length percentiles at birth and status for infants with (light gray bars, dotted line) and without (dark gray bars, solid line) LGR (A), and of BMI percentiles at birth and status for LGR infants with elevated (light gray bars, dotted line) versus low-normal (dark gray bars, solid line) BMI (B). The bars are constructed as histograms and the curves represent smoothed density curves. The third shade of gray noted is where the histogram bars overlap for each of the two groups. Note the leftward shift in length percentiles from...

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birth to status in infants with LGR (A), and the separation of BMI curves from birth to status in LGR infants (B) suggesting different patterns of weight gain relative to linear growth.
Figure 2.
Bar graph demonstrating the observed mean composite Bayley scores for infants without LGR (dark solid gray columns), with LGR (light solid gray columns), with LGR and elevated BMI (checkered gray columns) and with LGR and low-normal BMI (striped gray columns). The black bars represent the standard error. The # denotes a p-value < 0.05 between groups on unadjusted analysis. After multivariate analyses, only the observed differences in language scores between infants with and without LGR (*, Model 1, p<0.05) and LGR infants with high versus low-normal BMI (^, Model 2, p<0.05) remained significant.
Table 1.

Maternal and Infant Characteristics

| Variable                      | No LGR | LGR | Bivariate Analyses |
|-------------------------------|--------|-----|-------------------|
|                               | N = 315| All N = 912 | BMI at status ≥ 75th centile N = 353 (39%) | BMI at status < 75th centile N = 559 (61%) | p, no LGR vs. LGR | p, high vs. low BMI, LGR |
| Maternal race, n (%)          |        |     |                   |                   |                   |                   |
| African American              | 150 (48.4) | 396 (44.5) | 154 (45.2) | 242 (44.2) | 0.423 | 0.422 |
| Caucasian                     | 143 (46.1) | 448 (50.4) | 174 (51.0) | 274 (50.0) | 0.422 | 0.422 |
| Other                         | 17 (5.5) | 45 (5.1) | 13 (3.8) | 32 (5.8) | 0.422 | 0.422 |
| Maternal education, n (%)     |        |     |                   |                   |                   |                   |
| Less than high school         | 66 (21.0) | 170 (18.6) | 70 (19.8) | 100 (17.9) | 0.162 | 0.506 |
| High school grad              | 94 (29.8) | 305 (33.4) | 112 (31.7) | 193 (34.5) | 0.506 | 0.162 |
| College                       | 154 (49.0) | 422 (46.3) | 163 (46.2) | 259 (46.3) | 0.506 | 0.162 |
| Private insurance, n (%)      | 108 (34.3) | 307 (33.7) | 112 (31.7) | 195 (34.9) | 0.836 | 0.350 |
| PIH, n (%)                    | 32 (10.2) | 102 (11.2) | 42 (11.9) | 60 (10.7) | 0.836 | 0.350 |
| Maternal Diabetes, n (%)      | 16 (5.1) | 38 (4.2) | 14 (4.0) | 24 (4.3) | 0.524 | 0.866 |
| Chorioamnionitis, n (%)       | 163 (58.2) | 494 (54.7) | 194 (58.1) | 300 (56.5) | 0.781 | 0.672 |
| Male, n (%)                   | 153 (48.6) | 481 (52.7) | 197 (55.8) | 284 (50.8) | 0.214 | 0.153 |
| GA at birth, weeks            | 25.9 ± 1.2 | 25.6 ± 1.2 | 25.4 ± 1.1 | 25.7 ± 1.3 | <0.001 | 0.003 |
| BW, grams                     | 808 ± 155 | 776 ± 152 | 791 ± 157 | 766 ± 148 | <0.001 | 0.017 |
| SGA 1, n (%)                  | 28 (8.9) | 119 (13.0) | 33 (9.4) | 86 (15.4) | 0.056 | 0.009 |
| Length at birth, cm           | 32.4 ± 2.3 | 32.9 ± 2.4 | 32.8 ± 2.4 | 33.0 ± 2.5 | 0.001 | 0.365 |
| HC at birth, cm               | 23.1 ± 1.6 | 22.9 ± 1.6 | 22.9 ± 1.6 | 22.8 ± 1.6 | 0.016 | 0.419 |
| BMI at birth, kg/m²           | 7.7 ± 0.9 | 7.1 ± 0.8 | 7.3 ± 0.8 | 7.0 ± 0.8 | <0.001 | <0.001 |
| Variable                               | No LGR | LGR | Bivariate Analyses |
|---------------------------------------|--------|-----|--------------------|
|                                       | N = 315| All N = 912 | BMI at status ≥75\(^{th}\) centile N = 353 (39%) | BMI at status < 75\(^{th}\) centile N = 559 (61%) | \(p\), no LGR vs. LGR | \(p\), high vs. low BMI, LGR |
| DOL of first feed, median (IQR)       | 4 (2–5) | 3 (2–5) & 3 (2–5) | 4 (3–5) | 0.518 | 0.130 |
| DOL full feed, median (IQR)           | 20 (15–29) | 21 (15–32) | 20 (15–28) | 22 (16–35) | 0.330 | 0.001 |
| Days received parenteral nutrition, median (IQR) | 21 (15–30) | 23 (15–36) | 21 (14–32) | 24 (16–39) | 0.034 | 0.002 |

Data shown are mean ± SD unless otherwise noted.

\(^{1}\) SGA defined as birth weight z-score < −1 using Fenton growth curve
## Table 2.

### Infant Characteristics at Status

| Variable                  | No LGR N = 315 | All N = 912 | LGR BMI at status ≥ 75th centile N = 353 (39%) | LGR BMI at status < 75th centile N = 559 (61%) | Bivariate Analyses |
|---------------------------|----------------|-------------|------------------------------------------------|------------------------------------------------|-------------------|
| **Measures**              |                |             |                                                |                                                |                   |
| PMA, weeks                | 40.0 ± 2.6     | 40.0 ± 2.5  | 40.3 ± 2.4                                     | 39.8 ± 2.6                                     | 0.845             |
| Weight, grams             | 3209 ± 660     | 2922 ± 632  | 3288 ± 595                                     | 2692 ± 539                                     | <0.001            |
| Length, cm                | 49.1 ± 3.2     | 45.6 ± 3.2  | 45.6 ± 3.4                                     | 45.6 ± 3.1                                     | <0.001            |
| HC, cm                    | 34.2 ± 2.4     | 33.3 ± 2.3  | 34.1 ± 2.4                                     | 32.9 ± 2.0                                     | <0.001            |
| BMI, kg/m²                | 13.2 ± 2.4     | 13.9 ± 2.5  | 15.8 ± 2.0                                     | 12.8 ± 1.3                                     | <0.001            |
| **Outcomes**              |                |             |                                                |                                                |                   |
| Culture-positive sepsis, n (%) | 58 (18.4)    | 233 (25.6)  | 92 (26.1)                                      | 141 (25.2)                                     | 0.011             |
| NEC requiring surgery, n (%) | 5 (1.6)      | 17 (1.9)    | 3 (0.8)                                        | 14 (2.5)                                       | 1.000             |
| IVH (grade 3–4), n (%)    | 32 (10.2)      | 142 (15.6)  | 62 (17.6)                                      | 80 (14.4)                                      | 0.015             |
| PVL, n (%)                | 15 (4.8)       | 47 (5.2)    | 21 (6.0)                                       | 26 (4.7)                                       | 0.882             |
| Postnatal steroids, n (%) | 51 (18.3)      | 182 (22.5)  | 82 (26.9)                                      | 100 (19.9)                                     | 0.150             |
| BPD, n (%)                | 168 (53.5)     | 450 (49.5)  | 198 (56.2)                                     | 252 (45.2)                                     | 0.239             |
| ROP requiring surgery, n (%) | 13 (4.2)    | 80 (8.8)    | 23 (6.6)                                       | 57 (10.3)                                      | 0.006             |
| PDA requiring surgery, n (%) | 27 (8.6)    | 112 (12.3)  | 34 (9.6)                                       | 78 (14.0)                                      | 0.080             |

Data shown are mean ± SD unless otherwise noted.
## Table 3.

### Neurodevelopmental Outcomes at 22–26 Months

| Measure                          | No LGR | LGR       | Multivariate Analyses |
|----------------------------------|--------|-----------|-----------------------|
|                                  | N = 315| All N = 912| **p, no LGR vs. LGR** | **p, high vs. low BMI, LGR** |
| BMI at status ≥ 75th centile     |        | BMI at status < 75th centile | Model 1 | Model 2 | Model 1 | Model 2 |
| N = 353 (39%)                    |        | 355 (61%) | 0.367 | 0.538 | 0.247 | 0.111 |
| Moderate/severe cerebral palsy, n (%) | 17 (5.4) | 35 (9.9) | 0.047 | 0.588 | 0.058 | **0.030** |
| GMFCS level 2+, n (%)            | 23 (7.3) | 45 (12.8) | 0.348 | 0.992 | 0.115 | **0.037** |
| NDI, n (%)                       | 45 (14.3) | 73 (20.7) | 0.348 | 0.992 | 0.115 | **0.037** |

1. Model 1 variables: GA, birth weight, gender, SGA, maternal insurance, and center

2. Model 2 variables: All in Model 1 plus postnatal head sparing