Nonlinear Systems

Relaxation of Conditions for Convergence of Dynamic Regressor Extension and Mixing Procedure

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Abstract—A generalization of the dynamic regressor extension and mixing procedure is proposed, which, unlike the original procedure, first, guarantees a reduction of the unknown parameter identification error if the requirement of regressor semi-finite excitation is met, and second, it ensures exponential convergence of the regression function (regressand) tracking error to zero when the regressor is semi-persistently exciting with a rank one or higher.

Keywords: identification, linear regression, semi-finite excitation, semi-persistent excitation, parameter error, convergence, boundedness, monotonicity, singular value decomposition

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1. INTRODUCTION

In recent years, in the literature on adaptive control and identification theory, more than a hundred papers have been published (see the references, the ones therein and review [1]) devoted to development of methods to identify unknown time-invariant parameters of linear regression equations with improved properties both in terms of transient quality indexes and the necessary conditions for parameters estimates convergence to their true values. A considerable part of these studies is based on the Dynamic Regressor Extension and Mixing (DREM) procedure [2] and its analogs (integral modification I-DREM [3], procedures to generate new scalar excited regressor G+D and D+G [4, 5], scalar identification schemes with finite-time convergence [6], etc.).

The basic DREM procedure [2] consists of the regressor extension and mixing steps. In the first step, the initial linear regression, which regressor is usually a vector, is transformed into an extended one with a square regressor matrix using stable dynamic operators and special extension schemes [1, 7, 8]. In the second step, the obtained equation is multiplied with the extended regressor adjunct matrix to convert it into a set of scalar equations with the same scalar regressor.

In contrast to the well-known conventional gradient identifier [9], the DREM procedure [2]: (1) allows one to introduce a set of scalar estimation laws, each of which is responsible for identification of a certain unknown parameter, and the accuracy and convergence rate of such identification can be improved by adjustment of such laws scalar adaptive gains, and (2) relaxes the regressor persistent excitation requirement and guarantees asymptotic convergence of estimates to the true values if the scalar regressor is non-square integrable. Modified DREM procedures [3–6], in turn, relax this condition and ensure exponential or finite-time convergence of the parameter error to zero if the regressor is finitely or initially exciting.

However, as it has been analytically proved and experimentally demonstrated in [7, 8], for DREM like procedures [2–6] the condition of the regressor finite excitation is necessary to obtain a scalar regressor that is bounded away from zero, and therefore it is a convergence condition. If this requirement is not met in schemes [2–6], the unknown parameters identification error, as well as the regressand tracking error, cannot be reduced. At the same time, even when the condition of regression finite excitation is not satisfied, the classical gradient identifier [9] ensures the unknown parameters identification error reduction and an asymptotic convergence of the tracking error, which significantly narrows the applicability domain of the DREM-like procedures [2–6] in comparison with this approach.
Generally speaking, the condition of the regressor finite excitation is quite a weak requirement [10] and not satisfied in two main situations: (1) at least one element of the regressor is identically zero; (2) a linear dependence between the components of the regressor occurs [5].

It is proved in [10] that the state vector of a stationary plant in the Frobenius form is excited finitely over the initial time interval if the reference signal is non-differentiable at least at one point of such interval, which is true, for example, if the reference signal is a Heaviside function. However, practical experience makes it possible to conclude that for each specific identification problem and each specific parametrization there exist their own particular requirements, which are necessary to ensure the regressor finite excitation. Currently, no generalized formalized criteria accepted by the control community have been proposed to verify a priori that the regressor is finitely exciting for an arbitrary parametrization. Therefore, as far as the identification and adaptive control problems are concerned, it is necessary to apply only such identification procedures and algorithms that are capable of ensuring the reduction of the unknown parameter estimation error and the convergence of the tracking error even when the regressor finite excitation is not provided, which, in particular, motivates the development of a modified dynamic regressor extension and mixing procedure with a relaxed convergence condition.

Such a relaxed requirement could be, for example, a semi-finite excitation condition, which, in contrast to the finite excitation condition, is met as long as at least one of the regressor elements is non-zero, even in case of linear dependence between all the regressor components [11].

To date, two main approaches [12–14] have been proposed in the literature known to the authors that relax the convergence condition of the basic DREM procedure to the requirement of semi-finite excitation.

In [12], an identification law with switches has been proposed, in which the I-DREM-based law is used when the condition of finite excitation is satisfied, and the conventional gradient law is applied when the requirement of semi-finite excitation is met. The main disadvantage of this approach is that in the second case it ensures the unknown parameter identification quality that coincides with the conventional gradient identifier. In [13, 14], on the basis of the modified Gramm–Schmidt process, the algorithm to remove linearly dependent rows and columns from the extended regressor matrix has been developed, which allows one to reduce the problem of the unknown parameters identification to the problem of numerical solution of algebraic equations system in case the analytical dependence of the unknown parameters from each other is known. However, there are some hesitations that these equations can be solved when the unknown parameters are independent from each other and, consequently, the extension of the method from [13, 14] to the general case faces difficulties.

Thus, the problem to relax the convergence condition of the basic procedure of dynamic regressor extension and mixing is actual and does not have effective solutions up to date. Therefore, in this study a new step of regularization of the extended regressor is proposed to be added to the conventional DREM procedure to relax its convergence condition.

The aim of the regularization step is to, first, check the conditions that are necessary and sufficient to generate a scalar separated-from-zero regressor and, second, virtually change the matrix of the extended regressor when such conditions are violated. More specifically, in the regularization step we propose to apply the eigenvalue decomposition to the extended regressor obtained by the Kreisselmeier filter [1], which, because of such regressor symmetry and positive semi-definiteness, allows one to:

—verify that the condition of finite excitation of the extended regressor is met by analysis of its eigenvalues;
—following the ridge regression method [15, 16], substitute zero eigenvalues of the regressor with arbitrary constants.

When the semi-finite excitation condition is met, mixing of the extended regressor modified by the regularization allows one to obtain a new regression with a non-zero scalar regressor over the semi-finite excitation time interval. Such result is impossible without regularization. In this study it is shown that the identification law based on such regression coincides with the DREM-based one if the regressor finite excitation requirement is met and, in addition, if the necessary condition of semi-finite excitation and a number of sufficient conditions are satisfied, it ensures both the identification and tracking errors decrease.

The main result of this research is a dynamic regressor extension, regularization, and mixing procedure that relaxes the convergence condition of the basic DREM method.
The definitions from [3, 9–11, 17], which are used in axiomatic manner to state the problem and present the main result, are introduced.

**Definition 1.** The regressor $\varphi(t) \in R^n$ is persistently exciting ($\varphi(t) \in \text{PE}$), if $\forall t \geq t_0 \geq 0 \exists T > 0$ and $\alpha > 0$ such that the following holds

$$\lambda_{\min} \left\{ \int_{t}^{t+T} \varphi(\tau)\varphi^T(\tau)d\tau \right\} \geq \alpha,$$

(1.1)

where $\alpha > 0$ is the excitation level, $\lambda_{\min}()$ stands for the operator that returns the minimum eigenvalue of a matrix.

**Definition 2.** The regressor $\varphi(t) \in R^n$ is finitely exciting ($\varphi(t) \in \text{FE}$) over the time range $[t'_c; t_e] \subset [t_0; \infty)$, if there exist $t_e > t'_c \geq t_0 > 0$ and $\alpha > 0$ such that

$$\lambda_{\min} \left\{ \int_{t'_c}^{t_e} \varphi(\tau)\varphi^T(\tau)d\tau \right\} \geq \alpha.$$

(1.2)

**Definition 3.** The regressor $\varphi(t) \in R^n$ is semi-persistently exciting ($\varphi(t) \in \text{s-PE}$) with the time-invariant rank $0 < r < n$, if $\forall t > t_0 \geq 0 \exists T > 0$ and $0 < \underline{\alpha} < \overline{\alpha}$ such that $\forall i \in \{1, \ldots, r\}$ the inequality holds

$$\underline{\alpha} \leq \lambda_i \left\{ \int_{t'_c}^{t_e} \varphi(\tau)\varphi^T(\tau)d\tau \right\} \leq \overline{\alpha},$$

(1.3)

where $0 < \underline{\alpha} \leq \overline{\alpha}$ is a partial excitation level.

**Definition 4.** The regressor $\varphi(t) \in R^n$ is semi-finitely exciting ($\varphi(t) \in \text{s-FE}$) with time-invariant rank $0 < r < n$ over the time range $[t'_c; t_e] \subset [t_0; \infty)$, if there exists $t_e > t'_c$ and $0 < \underline{\alpha} < \overline{\alpha}$ such that $\forall i \in \{1, \ldots, r\}$ it holds that

$$\underline{\alpha} \leq \lambda_i \left\{ \int_{t'_c}^{t_e} \varphi(\tau)\varphi^T(\tau)d\tau \right\} \leq \overline{\alpha}.$$

(1.4)

The relations between the introduced regressor excitation types are specified as follows:

$$\varphi(t) \in \text{PE} \Rightarrow \left\{ \begin{array}{l} \varphi(t) \in \text{FE} \\ \varphi(t) \in \text{s-PE} \end{array} \right\} \Rightarrow \varphi(t) \in \text{s-FE}.$$

The requirements (1.1) and (1.2) impose constraints on all eigenvalues of the Gramm matrix, whereas (1.3) and (1.4) restrict only some of them. That is why the condition $\varphi(t) \in \text{s-FE}$ is the weakest one and, as far as limiting case is considered, is met when $r = 1$ if at least one of the regressor $\varphi(t)$ elements is non-zero over the time range $[t'_c; t_e] \subset [t_0; \infty)$.

An important role in the modern identification theory is played by the Kreisselmeyer filtering, which allows one to transform a vector regressor $\varphi(t) \in R^n$ into a matrix one $\varphi(t) \in R^{n\times n}$ without loss of excitation:

$$\forall t \geq t_0 \varphi(t) = -l\varphi(t) + \varphi(t)\varphi^T(t), \quad \varphi(t_0) = 0_{\text{max}},$$

(1.5)

where $l > 0$ is the Kreisselmeyer filter parameter.

The properties of the matrix regressor $\varphi(t) \in R^{n\times n}$ with respect to the conditions (1.1)–(1.4) on $\varphi(t) \in R^n$ are copied from [8, 11].

**Corollary 1.** $\varphi(t) \in \text{PE} \Leftrightarrow \forall t \geq kT \lambda_{\min}(t) > \mu$.

**Corollary 2.** $\varphi(t) \in \text{FE} \Leftrightarrow \forall t \in [t_0; t_0 + \delta] \subset [t'_c; t_e] \lambda_{\min}(t) > \mu$.

**Corollary 3.** $\varphi(t) \in \text{s-PE} \Leftrightarrow \forall t \geq kT \forall i \in \{1, \ldots, r\} \lambda_i(t) > \mu$. 


**Corollary 4.** \( \Phi(t) \in s\text{-FE} \Leftrightarrow \forall t \in [t_0; t_0 + \delta] \subset [t_*; t_*^+] \) \( \forall i \in \{1,\ldots,r\} \lambda_i(t) > \mu \).

Here, \( k \geq 1 \) is a positive integer number, \( \mu > 0 \) is a lower bound of the eigenvalue, \( \lambda_i(t) \) is the \( i \)-th eigenvalue of the regressor \( \varphi(t) \), \( \lambda_{\min}(t) = \min_{i \in \mathbb{N}^{n \times r}} \lambda_i(t) \) is the minimum separated-from-zero eigenvalue of the regressor \( \varphi(t) \), \( r = n - r \) is the rank deficiency.

The proofs of Corollary 1 and 2 are given in [8, 11], respectively, while the proofs of Corollary 3 and 4 can be obtained in the same way.

Based on the definition of the eigenvalue decomposition of the positive semi-definite time-invariant matrix from [17], the definition of the eigenvalue decomposition of the dynamic regressor \( \varphi(t) \in R^{n \times r} \) is introduced.

**Definition 5.** The eigenvalue decomposition of the regressor \( \varphi(t) \in R^{n \times r} \) with piecewise-constant rank \( r(t) \leq n \) is defined as follows:

\[
V^T(t)\varphi(t)V(t) = \begin{pmatrix} V_1^T(t) \\ V_2^T(t) \end{pmatrix} \varphi(t) \begin{pmatrix} V_1(t) \\ V_2(t) \end{pmatrix} = \Lambda(t) = \begin{bmatrix} \Lambda_1(t) & 0_{r(t) \times r(t)} \\ 0_{r(t) \times r(t)} & 0_{r(t) \times r(t)} \end{bmatrix},
\]

where \( V_1(t) \in R^{r(t) \times r(t)} \) stands for a time-varying orthonormal basis of \( \varphi(t) \) eigenspace, \( V_2(t) \) is a time-varying orthonormal basis of \( \varphi(t) \) nullspace, \( \lambda_1(t) \geq \lambda_2(t) \geq \ldots \geq \lambda_{r(t)}(t) > 0 \) denote nonzero eigenvalues of \( \varphi(t) \), \( 0_{r(t) \times r(t)} \in R^{r(t) \times r(t)} \) is a zero matrix, \( 0_{r(t) \times r(t)} \in R^{r(t) \times r(t)} \) stands for zero matrices of corresponding dimensions.

### 2. PROBLEM STATEMENT

The classical problem of the time-invariant parameters identification of a linear regression equation is considered:

\[
\forall t \geq t_0 \quad z(t) = \Phi(t)^T(t)\theta,
\]

where \( \Phi(t) \in R^n \), \( z(t) \in R \) are measurable regressor and function (regressand), \( \theta \in R^n \) is a vector of unknown time-invariant (\( \theta \equiv 0 \)) and bounded (\( \|\theta\| \leq \theta_{\max} \)) parameters.

It is assumed that the following assumption holds for \( \Phi(t) \).

**Assumption 1.** The regressor \( \Phi(t) \) is bounded: \( \|\Phi(t)\| \leq \Phi_{\max} \).

In general case, the above-stated requirement can be met with the help of multiplication of (2.1) by \( n_t = \frac{1}{1 + \Phi^T(t)\Phi(t)} \).

The aim is to derive the adaptive law to obtain the estimations \( \hat{\theta}(t) \in R^n \), which, when \( \Phi(t) \in s\text{-FE} \), ensures that:

\[
\|\hat{\theta}(t)\| \leq \beta \|\theta(t)\|, \quad \beta \in (0, 1),
\]

where \( \hat{\theta}(t) = \Phi(t)\hat{\theta} - z(t) \) is the tracking error, \( \hat{\theta}(t) = \hat{\theta}(t) - \theta \) stands for the parameter error.

The inequalities (2.2) mean the reduction of \( \hat{\theta}(t) \) and \( z(t) \), respectively, over the time range \( [t_*; t_*^+] \). The requirement \( \Phi(t) \in s\text{-FE} \) is the convergence condition of the desired adaptive law. The convergence is capability to reduce the initial values of the errors \( \hat{\theta}(t) \) and \( \hat{\theta}(t) \).
2.1. Gradient-Based Identification Law

The classical solution, which ensures that the goal (2.2) is achieved, is the gradient-based identification law:

\[ \dot{\hat{\theta}}(t) = -\Gamma \varphi(t)(\Psi^T(t)\hat{\theta}(t) - z(t)), \quad \Gamma = \Gamma^T > 0, \]  

which convergence is guaranteed when \( \varphi(t) \in s\text{-FE} \) and, in general, it ensures the following properties:

\begin{align*}
(a_1) \quad & \varphi(t) \in PE \iff \lim_{t \to \infty} \|\hat{\theta}(t)\| = 0 \text{(exp)} \quad \lim_{t \to \infty} \|\dot{\hat{\theta}}(t)\| = 0 \text{(exp)}; \\
(a_2) \quad & \lim_{t \to \infty} |\dot{\hat{\theta}}(t)| = 0; \\
(a_3) \quad & \lambda_{\text{min}}(\Gamma) = \lambda_{\text{max}}(\Gamma) \Rightarrow \|\hat{\theta}(t)\| \leq \|\hat{\theta}(t_0)\| \quad \forall t_a \geq t_0; \\
(a_4) \quad & \varphi(t) \in s\text{-FE} \Rightarrow \left\{ \begin{array}{l}
\|\hat{\theta}(t)\| \leq \|\hat{\theta}(t_0)\|, \\
\|\dot{\hat{\theta}}(t)\| \leq \|\dot{\hat{\theta}}(t_0)\|, \\
\end{array} \right. \beta \in (0; 1), \\
(a_5) \quad & \text{when } \varphi(t) \in PE \text{ there is an optimal value of } \Gamma \text{ that maximizes the rate of exponential convergence of the parameter error } \hat{\theta}(t) \text{ to zero. The change of any element of the matrix } \Gamma \text{ affects the transients quality of all } \hat{\theta}(t). 
\end{align*}

Despite ensuring some properties when \( \varphi(t) \in s\text{-FE} \) (a_4), the law (2.3) guarantees exponential convergence of \( \hat{\theta}(t) \) and \( \dot{\hat{\theta}}(t) \) to zero if the strict condition of the regressor persistent excitation (a_1) is met, provides monotonicity of the parameter error norm only (a_3), and each element of the arbitrary parameter \( \Gamma \) affects the transients quality of all errors (a_5).

To overcome the disadvantages of the law (2.3), in [2] a DREM procedure has been proposed, according to which, firstly, the regression (2.1) is processed using the regressor extension and mixing operations, and then, on the basis of the obtained new regression, the unknown parameter identification law is introduced. The synthesis procedure and properties of such a law are considered below.

2.2. Dynamic Regressor Extension and Mixing

In the step of extension the initial vector regressor \( \varphi(t) \in R^n \) is transformed into the matrix one \( \varphi(t) \in R^{n \times n} \) using, for an instance, the filter (1.5):

\[ \dot{\varphi}(t) = -\varphi(t) + \bar{\varphi}(t) \varphi^T(t), \quad \varphi(t_0) = 0_{n \times n}, \]  

\[ \dot{y}(t) = -y(t) + \bar{\varphi}(t)z(t), \quad y(t_0) = 0, \]  

where \( y(t) \in R^n \) is the extended regressand.

After filtering (2.4) the extended regression equation is obtained:

\[ y(t) = \varphi(t)\theta, \]  

which regressor, in accordance with Corollary 1–4, could be used to verify the fact that any of the conditions (1.1)–(1.4) is met.

In the mixing step, in accordance with [2], the matrix regressor \( \varphi(t) \in R^{n \times n} \) is transformed into scalar one \( \omega(t) \in R \) by way of multiplication of (2.5) by the adjoint matrix \( \text{adj}\{\varphi(t)\} \) and application of the property \( \text{adj}\{\varphi(t)\}\varphi(t) = \text{det}\{\varphi(t)\}I_{n \times n} \):

\[ Y(t) = \omega(t)\theta, \]  

\[ Y(t) := \text{adj}\{\varphi(t)\}y(t), \quad \omega(t) := \text{det}\{\varphi(t)\}, \]  

where \( Y(t) \in R^n \).

On the basis of the obtained \( n \) scalar equations (2.6) the following identification law is introduced according to [2]:

\[ \dot{\hat{\theta}}(t) = \hat{\theta}(t) = -\gamma_1 \omega(t)(\omega(t)\dot{\hat{\theta}}(t) - \omega(t)\dot{\theta}(t)) = -\gamma_1 \omega(t)\dot{\hat{\theta}}(t), \quad \gamma_1 > 0, \]  

where \( \gamma_1 > 0 \) is a positive constant.
which convergence condition is $\bar{\varphi}(t) \in \text{FE}$, and it ensures the following properties:

1. $\alpha(t) \notin L_2 \iff \lim_{t \to \infty} \| \bar{\theta}(t) \| = 0$,
2. $\alpha(t) \notin \text{PE} \iff \lim_{t \to \infty} \| \bar{\theta}(t) \| = 0 \text{ (exp)}$;
3. $\lim_{t \to \infty} \| \bar{\theta}(t) \| = 0 \Rightarrow \lim_{t \to \infty} \left[ \left( z(t) - \bar{\varphi}^T(t) \hat{\theta}(t) \right) \right] = 0 \text{ (certainty equivalence)}$;
4. $\left| \bar{\theta}(t_n) \right| \leq \left| \bar{\theta}(t_0) \right| \ \forall t_n \geq t_b$;
5. $\bar{\varphi}(t) \in \text{FE} \Rightarrow \left\{ \left| \bar{\theta}(t_n) \leq \beta \left| \bar{\theta}(t_0) \right| \right| \; \beta \in (0; 1), \left| \tilde{z}(t_n) \right| \leq \beta \left| \tilde{z}(t_0) \right| \right\}$.

$b_n$ when $\bar{\varphi}(t) \in \text{PE}$, the exponential convergence rate of the parameter error $\bar{\theta}_i(t)$ can be improved by increase of $\gamma_i$ and change of any element $\gamma_i$ affects only the transient quality of the respective $\hat{\theta}_i(t)$.

As follows from the comparison of the properties $(a_1) - (a_2)$ and $(b_1) - (b_2)$, the relaxed requirement of asymptotic convergence of the parameter error $(b_1)$, the monotonicity of the transients of each particular error $\bar{\theta}_i(t)$ $(b_2)$ as well as the fact that the transients quality of estimates for each particular $\bar{\theta}_i(t)$ $(b_3)$ can be adjusted with the help of $\gamma_i$ are the advantages of (2.7) compared to the gradient law (2.3). However, at the same time, the law (2.7) does not provide convergence to zero of the error $\tilde{z}(t)$ separately from the parameter error convergence $(b_4)$ and has a stricter convergence condition $(b_5)$.

Therefore, the main goal of this study is to develop an identification law that combines the positive properties of (2.3) and (2.7), which means that, when $\bar{\varphi}(t) \in \text{FE}$, the proposed law is required to have the properties $(b_1) - (b_4)$ of (2.7), while when $\bar{\varphi}(t) \in s\text{-FE}$—the property $(a_4)$ of the law (2.3), and in contrast to (2.7) it is required to ensure the convergence of the tracking error $\tilde{z}(t)$ separately from the parameter error convergence $(a_5)$.

3. MAIN RESULT

3.1. Dynamic Regularization of Extended Regressor

Following Definition 5, the regression equation (2.5) is rewritten as

$$y(t) = \varphi(t)\theta = \left[ V_1(t) \ V_2(t) \right] \Lambda(t) \begin{pmatrix} V_1^T(t) \\ V_2^T(t) \end{pmatrix} \theta = V(t)\Lambda(t)V^T(t)\theta. \quad (3.1)$$

It should be noted that, when rank $\{\varphi(t)\} = r(t) < n$, the matrix $\Lambda(t)$ contains $\Phi(t) > 0$ zeros on the main diagonal, and therefore $\alpha(t) = \det(\varphi(t)) \equiv 0 \Rightarrow \left| \bar{\theta}(t_n) \right| = \left| \bar{\theta}(t_0) \right|$. As a result, in order to make the regressor determinant $\varphi(t)$ be bounded away from zero when the regression with the scalar regressor (2.6) is obtained, the zeros of the main diagonal of the matrix $\Lambda(t)$ are to be virtually substituted with nonzero numbers $[15, 16]$. To achieve this, we introduce the matrix $\Xi(t)$, which being added to $\Lambda(t)$ allows one to obtain a full rank matrix:

$$\Xi(t) = \bar{\Lambda}(t) - \Lambda(t),$$

$$\bar{\Lambda}(t) \equiv \begin{cases} 0_{n \times r}, & \text{if } \text{diag}(\bar{\lambda}_1(t), \bar{\lambda}_2(t), \ldots, \bar{\lambda}_n(t)) = \varepsilon I_{n \times r}, \\
\text{diag}(\bar{\lambda}_1(t), \bar{\lambda}_2(t), \ldots, \bar{\lambda}_n(t)), & \text{otherwise}, \end{cases} \quad (3.2)$$

$$\bar{\lambda}_i(t) \equiv \begin{cases} \lambda_i(t), & \text{if } \lambda_i(t) \geq \varepsilon, \\
\varepsilon, & \text{if } \lambda_i(t) < \varepsilon, \end{cases} \quad i = 1, n$$

where $\bar{\Lambda}(t)$ is a new matrix of eigenvalues, $\varepsilon > 0$ stands for a parameter that defines the value of the virtual eigenvalues, $\varepsilon \equiv 0$ denotes the parameter that defines the amplitude of the eigenvalues of $\varphi(t)$, which are considered to be equivalently equal to zero in the presence of computation errors and external disturbances.
The expression \( \pm V(t) \Xi(t)V^T(t)\theta \) is added to (3.1) to obtain:

\[
y(t) = \phi(t)\theta = V(t)\Lambda(t)V^T(t)\theta \pm V(t)\Xi(t)V^T(t)\theta = V(t)\hat{\Lambda}(t)V^T(t)\theta - V(t)\Xi(t)V^T(t)\theta = \Phi(t)\theta - V(t)\Xi(t)V^T(t)\theta,
\]

(3.3)

where \( \Phi(t) \in R^{n \times n} \) is a new regressor with the eigenvalues \( \hat{\Lambda}(t) \).

The equation (3.3) is multiplied by the matrix \( \text{adj}\{\Phi(t)\} \), and then the following properties are applied:

\[
\text{adj}\{\Phi(t)\} = \text{det}\{\Phi(t)\}\Phi^{-1}(t), \quad \Phi^{-1}(t) = V(t)\hat{\Lambda}^{-1}(t)V^T(t),
\]

\[
\text{adj}\{\Phi(t)\}\Phi(t) = \text{det}\{\Phi(t)\}I_n,
\]

to obtain:

\[
\Upsilon(t) = \alpha(t)\theta - \alpha(t)V(t)\hat{\Lambda}^{-1}(t)\Xi(t)V^T(t)\theta = \alpha(t)\Theta(t),
\]

\[
\Upsilon(t) = \text{adj}\{\Phi(t)\}\gamma(t), \quad \alpha(t) = \text{det}\{\Phi(t)\},
\]

(3.4)

where \( \Theta(t) \in R^n \) is a vector of new unknown parameters, \( d(t) \in R^n \) is a disturbance, which causes the difference between \( \Theta(t) \) and \( \theta \).

The properties of the new regressor \( \omega(t) \) are presented in the following proposition.

**Proposition 1.** Let the matrix \( \hat{\Lambda}(t) \) be obtained using the equation (3.2) in case \( \bar{\varepsilon} = 0 \), then the following implications hold:

1. \( \bar{\gamma}(t) \in \text{PE} \iff \forall t \ni kT \quad \alpha(t) \equiv \lambda_{\min}^n(t) > \mu^n > 0 \).
2. \( \bar{\gamma}(t) \in \text{FE} \iff \forall t \in [t_0; t_0 + \delta] \subset [t^n_1; t^n_2] \quad \alpha(t) \equiv \lambda_{\min}^n(t) > \mu^n > 0 \).
3. \( \bar{\gamma}(t) \in \text{s-PE} \iff \forall t \ni kT \quad \alpha(t) \equiv \min\{\lambda_{\min}^n(t), \varepsilon^n\} > 0 \).
4. \( \bar{\gamma}(t) \in \text{s-FE} \iff \forall t \in [t_0; t_0 + \delta] \subset [t^n_1; t^n_2] \quad \alpha(t) \equiv \min\{\lambda_{\min}^n(t), \varepsilon^n\} > 0 \).

Proof of Proposition 1 is postponed to Appendix.

Using the regression (3.4) and the properties proved in Proposition 1, the identification law with normalization of the regressor excitation is introduced in accordance with [18]:

\[
\hat{\theta}(t) = -\gamma(t)\alpha(t)\hat{\theta}(t) - \Upsilon(t) = -\gamma(t)\omega(t)\hat{\theta}(t) - \gamma(t)\omega^2(t)d(t)
\]

\[
= -\gamma(t)\omega^2(t)\left(\hat{\theta}(t) - \Theta(t)\right) = \gamma(t)\omega^2(t)\hat{\theta}(t) \Rightarrow \hat{\theta}(t) = \Theta(t)
\]

\[
\gamma(t) = \begin{cases} 
\gamma_1, & \text{if } \alpha(t) \leq \min\{\lambda_{\min}^n(t), \varepsilon^n\} \\
\frac{\gamma_0}{\omega^2(t)}, & \text{otherwise},
\end{cases}
\]

(3.5)

where \( \gamma_0 > 0, \gamma_1 > 0 \) are arbitrary parameters of the identification law, \( \hat{\theta}(t) \in R^n \) is the error of the vector \( \Theta(t) \) identification.

Owing to the algorithm to form the matrix \( \Xi(t) \), the following theorem is valid for the law (3.5).

**Theorem 1.** Let Assumption 1 be met and \( \bar{\varepsilon} = 0 \), then:

1. \( \bar{\gamma}(t) \in \text{PE} \iff \forall t \ni \beta_1 \theta_{\max}, \beta_1 > 1 \),
2. the multiplication \( \gamma_0 \delta \) is such that \( 1 / \beta_1 + e^{-0.5\gamma_0 \delta} \in (0; 1) \),
3. \( \alpha(t) \in L_2 \iff \lim_{t \to \infty} \|\hat{\theta}(t)\| \leq \theta_{\max} \).
(4) \( \overline{\varphi}(t) \in s-\text{PE} \Leftrightarrow \lim \limits_{t \to \infty} \|\tilde{\Theta}(t)\| \leq \theta_{\text{max}} (\exp). \)

In this case the rate of exponential convergence can be directly adjusted by value of the parameter \( \gamma_0 \).

Proof of Theorem 1 is given in Appendix.

As follows from the results of Theorem 1, unfortunately, the law (3.5) is not capable of achieving the goal (2.2) if the values of \( \hat{\Theta}(t^*_r) \) are chosen arbitrarily, because in a set with a bound \( \hat{\Theta}(t^*_r) \) could become greater than \( \|\tilde{\Theta}(t^*_r)\| \), which is a disadvantage of the law (3.5) compared to the conventional gradient one (2.3). Therefore, the necessary condition for convergence of (3.5) is a semi-finite excitation of the regressor \( \overline{\varphi}(t) \in s-\text{FE} \), while the sufficient condition is that premises (2.1) and (2.2) are met. Here, it should also be noted that the choice \( \hat{\Theta}(t^*_r) = 0_n \) guarantees that the error \( \tilde{\Theta}(t) \) does not increase over the time range \([t^*_r, t_f]\). So it can be stated that the law (3.5) is quasi-convergent in terms of (2.2) when only the necessary condition \( \overline{\varphi}(t) \in s-\text{FE} \) is satisfied.

Thus, according to the proposed dynamic regressor regularization procedure (3.1)–(3.4), on the one hand, when \( \overline{\varphi}(t) \in \text{FE} \), the matrix \( \Lambda(t) \) is not added with \( \Xi(t) \) to form a full-rank matrix, and the law (3.5) reduces to (2.7), and on the other hand, when \( \overline{\varphi}(t) \in s-\text{FE} \), the matrix \( \Lambda(t) \) is added with \( \Xi(t) \) to form a full rank matrix, and, in contrast to (2.7), (3.5) ensures convergence in terms of (2.2) if the sufficient conditions are satisfied.

When the law (3.5) is applied, the global stability of the errors \( \tilde{z}(t) \) and \( \tilde{\Theta}(t) \) is analyzed by making different assumptions about the rank \( r(t) \) and the basis of the nullspace \( V_2(t) \). In Subsection 3.2 it is assumed that they are time-invariant, whereas in Subsection 3.3 they are considered to be piecewise-constant functions.

### 3.2. Time-Invariant Rank and Basis of Nullspace

The following assumption about the time-invariance of the rank and nullspace basis of the regressor \( \varphi(t) \in R_{\text{rnc}} \) is introduced.

**Assumption 2.** There exists the decomposition (1.6) with the time-invariant matrix \( V_2(t) \equiv V_2 \) of \( \varphi(t) \in R_{\text{rnc}} \) with constant rank \( r(t) \equiv r < n, \varphi(t) \equiv \varphi > 0 \).

Under Assumption 2, the disturbance \( d(t) \equiv d \) and the unknown parameters \( \Theta(t) \equiv \Theta \) are also time-invariant.

When the law (3.5) is applied and Assumption 2 is met, taking into account the results of Proposition 1, the properties of \( \tilde{z}(t) \) and \( \tilde{\Theta}(t) \) are analyzed in Theorem 2. In its first statement the unconditional properties are presented, in the second one the properties are shown that are guaranteed when the convergence condition is met, and in the third and fourth statements the asymptotic and exponential stability conditions are presented.

**Theorem 2.** When Assumptions 1 and 2 are met, the following holds:

I. \( \forall t \geq t_0 \) \( |\tilde{\Theta}(t) - \tilde{\Theta}_r| \leq |\tilde{\Theta}_r| \) \( \forall t_r \geq t_0 \).

II. \( \overline{\varphi}(t) \in s-\text{FE} \) \( \Rightarrow \begin{cases} \|\tilde{\Theta}(t)\| \leq \beta \|\tilde{\Theta}(t^*_r)\| \vspace{0.1cm} \\ \|\tilde{z}(t)\| \leq \beta \|\tilde{z}(t^*_r)\| \end{cases} \)

III. \( \omega(t) \not\in L_2 \) \( \Rightarrow \lim \limits_{t \to \infty} \|\tilde{\Theta}(t)\| = 0 \vspace{0.1cm} \lim \limits_{t \to \infty} \|\tilde{z}(t)\| = 0 \vspace{0.1cm} \)

IV. \( \overline{\varphi}(t) \in s-\text{PE} \) \( \Rightarrow \lim \limits_{t \to \infty} \|\tilde{\Theta}(t)\| = 0 (\exp), \vspace{0.1cm} \lim \limits_{t \to \infty} \|\tilde{z}(t)\| = 0 (\exp) \).

In this case the rate of exponential convergence can be directly adjusted by value of the parameter \( \gamma_0 \).

Proof of Theorem 2 is given in Appendix.

**Remark 1.** The asymptotic stability condition \( \omega(t) \not\in L_2 \) is strictly weaker than the exponential one \( \overline{\varphi}(t) \in s-\text{PE} \), as, for example, there exists the regressor \( \omega(t) = e^{\pi t} \lambda(t) \), \( \lambda(t) = \frac{1}{\sqrt{1 + t}} \), such that \( \omega(t) \not\in L_2 \).
and \( \overline{\varphi}(t) \in s\text{-PE} \) because \( \exists \mu > 0 \ \forall t \geq t_0 \ \lambda_1(t) > \mu \), which contradicts Corollary 3. Therefore, when Assumption 2 is met, the weakest convergence condition of the law (3.5) to ensure convergence of the errors \( \overline{\Theta}(t), \overline{\vartheta}(t) \) to zero, and \( \theta(t) \) to the set \( \theta_{\text{max}} \) is the non-square integrability of the multiplication of \( r \) eigenvalues of \( \varphi(t) \).

### 3.3. Piecewise-Constant Rank and Nullspace Basis

The requirements of Assumption 2 are restrictive, and, as far as practical scenarios are concerned, both the rank and nullspace basis of the regressor usually change their values in piecewise-constant manner. Therefore, the properties of the law (3.5) are analyzed under the assumption that the rank and nullspace basis of \( \varphi(t) \) are defined as piecewise-constant functions.

**Assumption 3.** The rank of \( \varphi(t) \) is a piecewise-constant function, and there exists its decomposition (1.6) with the piecewise-constant matrix \( V_2(t) \):

\[
\forall t \geq t_0 \quad r(t) = \sum_{j_r=1}^{m_r} \Delta_{j_r} h(t - t_{j_r}), \quad V_2(t) = \sum_{j_r=1}^{m_r} \Delta_{j_r} h(t - t_{j_r}),
\]

where \( t_{j_r} \) is a time instant of rank change, \( \Delta_{j_r} \) denotes the amplitude of rank change at time instant \( t_{j_r} \), \( t_{j_r} \) stands for the time instant of change of the nullspace basis \( V_2(t) \), \( \Delta_{j_r} \in \mathbb{R}^{o_r(t)} \) is the amplitude of \( V_2(t) \) change, \( h(t - t_{j_r}) \), \( h(t - t_{j_r}) \) are unit step functions.

When (3.6) is met, the disturbance \( d(t) \) and unknown parameters \( \Theta(t) \) are piecewise-constant functions:

\[
d(t) = \sum_{j=1}^{m} \Delta_j h(t - t_j), \quad \dot{d}(t) = \sum_{j=1}^{m} \Delta_j \delta(t - t_j), \quad \Theta(t) = \theta - \sum_{j=1}^{m} \Delta_j h(t - t_j),
\]

where \( t_j \in \{t_{j_r} : j_r \in \mathbb{N}, j_r \in \mathbb{N} \} \) are time instants of \( d(t) \) change, \( \delta(t - t_j) \) is a Dirac function, \( \|\Delta_j\| \leq \Delta_{\text{max}} \) is a bounded value of the disturbance amplitude change.

Taking into consideration proved Proposition 1, the properties ensured by the law (3.5) when Assumptions 1 and 3 are met are stated in the following theorem.

**Theorem 3.** Let the premises of Assumptions 1 and 3 hold and \( \overline{\varphi}(t) \in s\text{-PE} \) with the rank \( r(t) \geq 1 \), then:

\[
\forall t \geq kT \quad \left\| \overline{\Theta}(t) \right\| \leq a(t) e^{-\gamma(t-kT)} \left\| \overline{\Theta}(kT) \right\|, \quad \left\| \overline{\vartheta}(t) \right\| \leq a(t) e^{-\gamma(t-kT)} \left\| \overline{\vartheta}(kT) \right\|,
\]

where \( \{a(t_0), a(t_1), \ldots, a(t_j), \ldots\} \) is a numerical sequence.

Moreover, when \( \exists a_{\text{max}} \forall t_j \geq t_0 \ a(t_j) \leq a_{\text{max}} \), then \( \overline{\Theta}(t) \) and \( \overline{\vartheta}(t) \) are exponentially stable:

\[
\lim_{t \to \infty} \left\| \overline{\vartheta}(t) \right\| = 0 \ (\text{exp}), \quad \lim_{t \to \infty} \left\| \overline{\Theta}(t) \right\| = 0 \ (\text{exp}).
\]

Proof of Theorem 3 and the definition of \( a(t_j) \) are presented in Appendix.

On the one hand, the results of Theorem 3 show the robustness of the law (3.5) to variations of the rank and nullspace basis of the regressor \( \varphi(t) \) in the sense of exponential recovery of equilibrium points of the errors \( \overline{\Theta}(t) \) and \( \overline{\vartheta}(t) \), and on the other hand, describe necessary and sufficient conditions of such errors exponential convergence to zero. These conditions are the regressor semi-persistent excitation with rank not less than one and the fact that the inequalities \( a(t_j) \leq a_{\text{max}} \) hold for all \( t_j \geq t_0 \).

However, Theorem 3 does not provide a constructive description of the requirements for \( a(t_j) \) or \( \Delta_j \), which, being met for all \( t_j \geq t_0 \), guarantee \( a(t_j) \leq a_{\text{max}} \) and hence exponential stability of the errors \( \overline{\Theta}(t) \) and \( \overline{\vartheta}(t) \) when the rank or nullspace basis are piecewise-constant functions.

In the following corollary, we introduce two additional conditions, under which for all \( t_j \geq t_0 \) it is ensured that the inequality \( a(t_j) \leq a_{\text{max}} \) holds.
Corollary 5. Let the premises of Theorem 3 be met and additionally one of the following conditions also hold:

1) \( j \leq j_{\text{max}} < \infty \);
2) \( \Delta_{\text{max}} \leq c(t_{j})e^{-\gamma(j-kT)}, \forall j \in \mathbb{N}, c(t_{j}) \equiv c(t_{j+1}) > 0. \)

Then there exists \( a_{\text{max}} \) such that \( \forall t_{j} \geq t_{0} a(t_{j}) \leq a_{\text{max}}. \)

Proof of Corollary 5 is given in Appendix.

According to the results of Corollary 5, the condition \( a(t_{j}) \leq a_{\text{max}} \) is met when the norm of the parameter change value \( \Delta_{\text{max}} \) is upper bounded by a decreasing sequence, or when a number of regressor null-space base/rank switches \( j \) is finite.

3.4. Conditions of Partial Identifiability

Considering the identification problems, the main aim is to ensure the convergence of the parameter error \( \hat{\theta}(t) \). Therefore, in addition to the results of Subsections 3.2 and 3.3, the conditions are defined under which the elements of the vector of new unknown parameters \( \Theta(t) \) partially or completely coincide with the elements of the original vector \( \theta \).

The analysis of the parameters \( \Theta(t) \) properties are written as a proposition.

Proposition 2. Let the matrix \( \bar{\Lambda}(t) \) be obtained with the help of (3.2) when \( \varepsilon = 0 \), then:

1) \( \bar{\varphi}(t) \in \text{FE}/\bar{\varphi}(t) \in \text{PE} \Rightarrow \Theta(t) = 0 \);

2) if Assumption 2 and the following conditions are met:
\[
\sum_{i=1}^{n-p} \omega_{i} \varphi_{i}(t) + \sum_{j=n-p+1}^{n} \omega_{j} \varphi_{j}(t) = 0, \quad w_{i} \neq 0, \quad w_{j} = 0,
\]

then \( \exists M \subset \{1, \ldots, n\}, |M| = p, \forall i \in M, \Theta_{i} = \theta_{i}. \)

Proof of Proposition 2 is presented in Appendix.\(^1\)

Thus, according to Proposition 2, the conditions of partial identifiability of parameters \( \theta \) are: (1) Assumption 2 is met, (2) the regressor \( \bar{\varphi}(t) \) is semi-persistently exciting, (3) \( n \) columns of the regressor \( \varphi(t) \) are linearly independent, (4) the identification problem dimension is \( n > 2 \). Combining the results of Theorem 1 and Proposition 2, a corollary is obtained that describes the convergence conditions for a part of parameter errors \( \hat{\theta}_{i}(t) \).

Corollary 6. Let Assumptions 1, 2 and the following conditions be met:
\[
\sum_{i=1}^{n-p} \omega_{i} \varphi_{i}(t) + \sum_{j=n-p+1}^{n} \omega_{j} \varphi_{j}(t) = 0, \quad w_{i} \neq 0, \quad w_{j} = 0, \quad n > 2.
\]

Then:

(a) \( \bar{\varphi}(t) \in \text{FE} \iff \forall i \in M \) \[
\left\{ \begin{array}{l}
|\hat{\theta}_{i}(t_{a})| \leq \beta |\hat{\theta}_{i}(t_{a}^{r})|,
|\hat{\theta}_{i}(t_{b})| \leq |\hat{\theta}_{i}(t_{a})| \quad \forall t_{a} \geq t_{b};
\end{array} \right.
\]

(b) \( \omega(t) \in L_{2} \iff \forall i \in M \) \[
\left\{ \begin{array}{l}
\lim_{t \to \infty} |\hat{\theta}_{i}(t)| = 0,
|\hat{\theta}_{i}(t_{a})| \leq |\hat{\theta}_{i}(t_{a})| \quad \forall t_{a} \geq t_{b};
\end{array} \right.
\]

(c) \( \bar{\varphi}(t) \in \text{PE} \iff \forall i \in M \) \[
\left\{ \begin{array}{l}
\lim_{t \to \infty} |\hat{\theta}_{i}(t)| = 0 \text{(exp)},
|\hat{\theta}_{i}(t_{a})| \leq |\hat{\theta}_{i}(t_{a})| \quad \forall t_{a} \geq t_{b};
\end{array} \right.
\]

Corollary 6 is obtained by combining the consistent premises and results of Theorem 1 and Statement 2.

\(^1\) In statement (2) of Proposition 2, without loss of generality, it is assumed that the first \( n - p \) columns of the regressor \( \varphi(t) = [\varphi_{1}(t) \ldots \varphi_{i}(t) \ldots \varphi_{n}(t)] \) are linearly dependent (in case \( \varphi(t) > 0 \) such form can always be obtained by columns permutation).
Remark 2. It is worth noting the existence of regressors $\phi(t)$ that do not satisfy the requirements of Proposition 2, but still ensure the existence of zero elements in the vector $d$ and allow one to identify some of the original unknown parameters $\theta$. For such regressors, the fact that some elements of $d$ are zero is not caused by the existence of zero rows/columns in the product $V_2^T V_2$ (see the proof of Proposition 2), but by the equality to zero of the elements of the product $V_2^T V_2 \theta$ (due to orthogonality of $V_2$ and $\theta$).

For an instance, if $\phi(t) = \begin{bmatrix} 1 & -1 \\ -1 & 1 \end{bmatrix}$, $\theta = \theta \begin{bmatrix} -1 & 1 \end{bmatrix}$, $\theta \neq 0$, then the premises of Proposition 2 do not hold, but $d = 0_n$, $\Theta = \theta$.

Remark 3. From the practical point of view, it is important not only to prove that some elements of the parameter vector $\Theta$ coincide with the elements of $\theta$ under some conditions, but also to indicate their positions in such vector. For this purpose, the indices of the zero rows of the basis $V_2$ can be used as such indicators if the premises of statement 2 of Proposition 2 are satisfied.

Remark 4. Under Assumption 3, the results of statement 2 of Proposition 2 are true locally over the time intervals when the regressor rank and nullspace basis are time-invariant. Hence, when the rank $r(t)$ changes its value, different number $p$ of elements of the vector $\theta$ can be identified over different time ranges $[t_{j-1}; t_j]$ and $[t_j; t_{j+1}]$, and when the regressor nullspace basis changes its value, different elements of vector $\theta$ can be identified over different time intervals $[t_{j-1}; t_j]$ and $[t_j; t_{j+1}]$.

4. MATHEMATICAL MODELLING

The DREM identification law with regularization (3.5) has been compared with the classical gradient (2.3) and DREM without regularization (2.7) ones in Matlab/Simulink. The simulation was conducted using numerical integration by the Euler method with a fixed discretization step $\tau_s = 10^{-4}$ s.

Subsections 4.1 and 4.2 presents the obtained simulation results under Assumptions 2 and 3, respectively.

4.1. Time-Invariant Rank and Nullspace Basis of Regressor

The regression equation (2.1) was defined as:

$$z(t) = \bar{\phi}^T(t) \theta = \begin{bmatrix} -2e^{-t} \cos(t) & e^{-t} \cos(t) \end{bmatrix} \theta = \begin{bmatrix} 4 & -8 \\ 12 \end{bmatrix}. \tag{4.1.1}$$

The parameters of the filter (2.4), algorithm of the eigenvalue virtual substitution (3.2) and identification laws (3.5), (2.3) were set as:

$$l = 100, \quad \varepsilon = 0.4, \quad \bar{\varepsilon} = 10^{-10}, \quad \gamma_0 = 5, \quad \gamma_1 = 1, \quad \Gamma = 5I_2. \tag{4.1.2}$$

In order to provide the same convergence rate for the laws (3.5) and (2.7), the adaptive gain $\gamma$ of the law (2.7) was defined similarly to (3.5), following the method of the regressor excitation normalization [18]:

$$\gamma(t) = \begin{cases} \gamma_1, & \text{if } \omega(t) \leq \min \{ \lambda_{\min}(t), \varepsilon \} \\ \frac{\gamma_0}{\omega(t)}, & \text{otherwise.} \end{cases} \tag{4.1.3}$$

First of all, it was shown that the convergence conditions of the laws (2.3), (2.7), and (3.5) were met. Figure 1 presents the behavior of the disturbance $d$ and the rank of regressor $\phi(t)$ in the course of the experiment.

As follows from the definition of the regressor $\bar{\phi}(t)$, Fig. 1a, the convergence conditions ($\bar{\phi}(t) \in$ s-FE) of laws (2.3) and (3.5) were met for all $t \geq 0$, whereas the convergence condition ($\bar{\phi}(t) \in$ FE) of the law (2.7) was not satisfied, so the simulation results are given only for the algorithms (3.5) and (2.3). It followed from Figs. 1a and 1b, that Assumption 2 was met, and, consequently, since $\bar{\phi}(t) \in$ s-FE, the law (3.5) guaranteed the errors $\tilde{\Theta}(t)$, $\tilde{\xi}(t)$ reduction in the course of the experiment. Moreover, as
Assumption 2 was satisfied, $d_2 = 0$ and $r = 2$, then the law (3.5) additionally ensured that the error $\varepsilon(t)$ decreased.

Firstly, it was set that $\theta_0 = [0 \ 0 \ 0]^T$, which meant that, according to Theorem 1, the law (3.5) was quasi-convergent (the reduction of $\|\hat{\varepsilon}(t)\|$ was guaranteed, as well as the lack of growth of $\|\hat{\varepsilon}(t)\|$ over the time range $[0; 1]$).

Figure 2 presents the transients of errors $\hat{\varepsilon}_i(t)$ of the laws (3.5)—(a) and (2.3)—(b).

The obtained transients indicate the advantages of (3.5) over (2.7) and the classical gradient (2.3) identification laws. In particular, unlike (2.7), the law (3.5) reduced the a priori values of the errors $\hat{\varepsilon}_i(t)$ and, unlike (2.3), ensured the transients of first-order type and monotonic exponential convergence of the error $\hat{\varepsilon}_i(t)$ to zero. The monotonicity of $\hat{\varepsilon}_i(t)$ can be explained by the fact that the condition $\theta_i \leq \Theta_1$, $\hat{\varepsilon}_i(t_0) > \Theta_1$ was met in the course of the experiment, which was a particular favorable situation.

Figure 3a shows a comparison of the error $\varepsilon(t)$ curves of the laws (3.5) and (2.3), while Fig. 3b presents the transients of the error $\hat{\theta}_i(t)$ when the law (3.5) was applied.

Figure 3a confirms that $\varepsilon(t)$ was reduced over the time range $[0; 1]$ when the law (3.5) was applied, Fig. 3b demonstrates the monotonicity of the error $\hat{\theta}_i(t)$ $\forall i \in 1, n$, which was proved analytically in Theorem 2.
Figure 4 shows the behavior of obtained with the help of the laws (3.5) and (2.3) under different initial conditions (for all initial conditions the law (3.5) was convergent or quasi-convergent).

The transients in Fig. 4 confirm the exponential convergence of the error \( \| \hat{\theta}(t) \| \) to a set with the bound \( \theta_{\max} \) proved in Theorem 1.

Then it was set that \( \theta_0 = [0 \ 10 \ 14]^T \), which did not satisfy the convergence conditions from Theorem 1 since \( \| \hat{\theta}(t_r^+) \| \approx 4.9 \) and \( \theta_{\max} \approx 15 \). Figure 5 shows the behavior of \( \| \hat{\theta}(t) \| \) under such choice of the initial conditions when the laws (3.5) and (2.3) were used.

The transients of \( \| \hat{\theta}(t) \| \) shown in Fig. 5 validated the conclusions made in Theorem 1. Indeed, when \( \| \hat{\theta}(t_r^+) \| < \theta_{\max} \) the convergence condition of the law (3.5) was not met, and, consequently, the error norm \( \| \hat{\theta}(t) \| \) could become greater than \( \| \hat{\theta}(t_r^+) \| \), and it was not ensured that all conditions of (2.2) were met.

Thus, the conducted numerical experiments fully confirmed the properties of the law (3.5) described within Theorems 1—2, Proposition 2, and Corollary 6 when \( \varphi(t) \in s\text{-FE} \) and Assumption 2 was met.

4.2. Piecewise-Constant Rank and Nullspace Basis of Regressor

4.2.1. First experiment. The regression equation (2.1) was defined as follows:

\[
\begin{align*}
\bar{z}(t) &= \bar{\varphi}^T(t) \bar{\theta} = [\bar{\varphi}_1(t) \ \bar{\varphi}_2(t) \ \bar{\varphi}_3(t)] \begin{bmatrix} 4 \\ -8 \\ 12 \end{bmatrix}, \\
\bar{\varphi}_1(t) &= 2 \sin(t), \quad 0 \leq t \leq 5 \\
\bar{\varphi}_2(t) &= \begin{cases} 2 \sin(t), & 5 < t \leq 15 \\ 4, & t > 15 \end{cases}, \\
\bar{\varphi}_3(t) &= \begin{cases} \sin(t), & 0 \leq t \leq 10 \\ \sin(50t), & 10 < t \leq 15 \\ \sin(t), & t > 15 \end{cases}.
\end{align*}
\]
The parameters of the filter (2.4), algorithm of the eigenvalue virtual substitution (3.2) and identification laws (2.3), (3.5) were set as:

\[ l = 100, \quad \varepsilon = 0.4, \quad \varepsilon = 10^{-10}, \quad \gamma_0 = 5, \quad \gamma_1 = 1, \quad \Gamma = I_3. \]  

In order to provide the same convergence rate for the laws (3.5) and (2.7), the adaptive gain \( \gamma \) of the law (2.7) was defined similarly to (3.5), following the method of the regressor excitation normalization [18]:

\[ \gamma(t) = \begin{cases} \gamma_1, & \text{if } \alpha(t) \leq \min(\lambda_{\min}(t), \varepsilon), \\ \frac{\gamma_0}{\omega(t)}, & \text{otherwise}. \end{cases} \]

(4.2.3)

First of all, it was shown that the convergence conditions of the laws (2.3), (2.7), and (3.5) were met. Figure 6 presents the behavior of the disturbance \( d(t) \) and rank of the regressor \( \varphi(t) \) in the course of the experiment.

As follows from Fig. 6a and Corollaries 2 and 4, the necessary condition of the convergence of (3.5) was met for all \( t \geq 0 \), while the convergence condition of (2.7) was satisfied only over the time range \( t \in [10; 15.34] \). According to Fig. 6 the number of parameter switches was finite \( j < \infty \) and \( r(t) < 1 \), and then, according to the results of Theorem 3 and Corollary 5, all necessary and sufficient conditions of exponential convergence of errors \( \tilde{z}(t) \) and \( \tilde{\Theta}(t) \) to zero were satisfied for (3.5). Moreover, since \( \forall t \in [5; 10] d_i(t) = 0 \), the partial identifiability conditions from Proposition 2 were also met over the time range \( [5; 10] \).

Having verified that the convergence conditions were met, the experiments were conducted using the algorithms (3.5), (2.7), and (2.3) under different initial conditions.

Firstly, it was set that \( \theta_0 = [0 \ 5 \ 10]^T \), which, according to Theorem 1, ensured that the necessary conditions of convergence of the law (3.5) were met:

\[ \beta_i = \|\hat{\Theta}(t_i^*)\| \|\theta\|^{-1} = \frac{18}{15} = 1.2; \quad \frac{1}{\beta_i} + e^{-\gamma \delta} = \frac{1}{1.2} + e^{-5.5} = 0.833 \in (0; 1). \]

Figure 7 depicts the transients of \( \hat{\theta}_i(t) \) for (3.5)—(a), (2.7)—(b), and (2.3)—(c).

The obtained transients confirmed the theoretical conclusions made in Remark 4. Indeed, if the conditions of the second statement of Proposition 2 were met over \([5; 10]\), then the law (3.5), in contrast to (2.7) and (2.3), ensured monotonicity for one element of the vector \( \hat{\Theta}(t) \). Comparing the quality of the transients, the advantages of the law (3.5) over (2.7) and (2.3) are seen. As for (2.3), the law (3.5) guaranteed the first-order type transient of \( \hat{\theta}_i(t) \) \( \forall i \in \overline{1,n} \). In comparison with (2.7), the law (3.5) converged not
only over the time range $[10; 15.34]$, but for all $t \geq 0$, and ensured that one element of the vector $\hat{\theta}(t)$ decreased to zero over $[5; 10]$.

Figure 8 presents the transients of $z(t)$ for the control systems based on the laws (3.5)—(a), (2.7)—(b), and (2.3)—(c).

The transients that are depicted in Fig. 8 validate that the tracking error $\hat{\theta}(t)$ recovered exponentially to its equilibrium, as it was is proved in Theorem 3, when $\bar{\theta}(t) \in s$-PE and Assumption 3 was met.
Figure 9 presents the behavior of the norm of \( \tilde{\Theta}(t) \). Having analyzed Fig. 9, it was concluded that the parameter error \( \tilde{\Theta}(t) \) recovered to its equilibrium point when \( \tilde{\Theta}(t) \in \text{s-PE} \) and Assumption 3 was met, which validated the conclusions made in Theorem 3.

As the number of the rank switches was finite \( j_{\max} < \infty \), then, according to Theorem 3 and Corollary 5, exponential recovery of \( \tilde{\Theta}(t) \) and \( \tilde{\Theta}(t) \) to their respective equilibrium points was equivalent to exponential stability.

Figure 10 presents transients of \( \| \tilde{\Theta}(t) \| \) when the laws (3.5), (2.7), and (2.3) were applied.

The transients of \( \| \tilde{\Theta}(t) \| \) obtained with the help of the law (3.5) confirmed the conclusions made in Theorem 1. The goal (2.2) was achieved when \( \tilde{\Theta}(t) \in \text{s-FE} \) and sufficient conditions were met, and \( \tilde{\Theta}(t) \) did exponentially converge to the set with the bound \( \theta_{\text{max}} \), while such properties were ensured by (2.3) only for all \( t \geq 5 \), and by (2.7)—only for \( \tilde{\Theta}(t) \in \text{FE} \).

Then it was set that \( \theta_0 = [0 \ -10 \ 14]^T \), which did not meet the sufficient convergence conditions of Theorem 1 since \( \| \tilde{\Theta}(t) \| \approx 4.9 \) and \( \theta_{\text{max}} = \| \theta \| \approx 15 \). Figure 11 shows the transients of \( \| \tilde{\Theta}(t) \| \) obtained under such choice of initial conditions when the laws (3.5), (2.7), and (2.3) were applied.
The simulation results shown in Fig. 11 follows the results of Theorem 1. Indeed, when $\|\tilde{\theta}(t')\| < \theta_{\text{max}}$, the law (3.5) did not converge (when $\|\tilde{\theta}(t')\| = \theta_{\text{max}}$, it was quasi-convergent), and the error norm $\|\tilde{\theta}(t)\|$ could become greater than $\|\tilde{\theta}(t')\|$.

4.2.2. Second experiment. The regression equation (2.1) was defined as:

$$z(t) = \tilde{\phi}^T(t)\theta = [\tilde{\phi}_1(t) \tilde{\phi}_2(t) \tilde{\phi}_3(t)]^T = [-8 12] \begin{bmatrix} 4 \\ 12 \end{bmatrix},$$

$$\tilde{\phi}_1(t) = \begin{cases} -2e^{-t} \cos(t), & 0 \leq t \leq 1, \\ e^{-t}, & 1 < t \leq 2, \\ e^t \cos(t), & t > 2; \end{cases}$$

$$\tilde{\phi}_2(t) = \begin{cases} e^t \cos(t), & 0 \leq t \leq 1, \\ -2e^{-t} \cos(t), & 1 < t \leq 2, \\ e^t + 0.1, & t > 2; \end{cases}$$

$$\tilde{\phi}_3(t) = \begin{cases} e^{-t}, & 0 \leq t \leq 1, \\ e^{-t} \cos(t), & 1 < t \leq 2, \\ -2e^{-t} \cos(t), & t > 2. \end{cases}$$

The parameters of the filter (2.4), algorithm of the eigenvalue virtual substitution (3.2) and identification laws (2.3), (3.5) were set as:

$$\gamma = 1, \quad \gamma_0 = 5, \quad \gamma_1 = 1, \quad \Gamma = I_3.$$  

In order to provide the same convergence rate for the laws (3.5) and (2.7), the adaptive gain $\gamma$ of the law (2.7) was defined similarly to (3.5), following the method of the regressor excitation normalization [18]:

$$\gamma(t) = \begin{cases} \gamma_0 & \text{if } \omega(t) \leq \min\{\lambda_{\text{min}}^*(t), \epsilon^\epsilon\}, \\ \gamma_1 \frac{\omega(t)}{\omega^2(t)} & \text{otherwise}. \end{cases}$$

First of all, it was shown that the convergence conditions of the laws (2.3), (2.7), and (3.5) were met. Figure 12 presents the behavior of the disturbance $d(t)$ and regressor $\varphi(t)$ rank in the course of the experiment.

The time ranges $[1; 1.165]$ and $[2; 2.14]$, at which rank $\{\varphi(t)\} = 3$, were substantially shorter than the time intervals, when rank $\{\varphi(t)\} = 2$. Therefore, unlike the experiment in Subsection 4.2.1, in this one the rank of the regressor was time-invariant almost everywhere. The rank differed from two when $[1; 1.165]$ and $[2; 2.14]$ as the filter (2.4) mixed information about regressors with different bases. Considering (2.7), the convergence condition was satisfied over $[1; 1.165]$ and $[2; 2.14]$ due to the mixing effect.

In turn, for the law (3.5) the necessary condition of convergence was satisfied for all $t \geq 0$. According to Fig. 12, the number of parameter switches was finite $j \leq j_{\text{max}} < \infty$ and $r \geq 1$, and then, by
Assumption 3 and the results of Theorem 3 and Corollary 5, for (3.5) all necessary and sufficient conditions of exponential convergence of the errors \( \tilde{z}(t) \) and \( \tilde{\Theta}(t) \) to zero were satisfied. Moreover, since

\[
\forall t \in [0; 1] \quad d_3(t) = 0, \quad \forall t \in [1; 2] \quad d_4(t) = 0, \quad \forall t \in [2; 3] \quad d_5(t) = 0,
\]

then the conditions of partial identifiability described in Proposition 2 were also met in the course of the experiment.

Having verified that the convergence conditions were met, the experiments were conducted using the algorithms (3.5), (2.7), and (2.3) under different initial conditions.

Firstly, it was set that \( \theta_0 = [0 -10 14]^T \), so the convergence conditions from Theorem 1 were not met since \( \|\theta(t^*)\| \approx 4.9 \), while \( \theta_{max} = \|\theta\| \approx 15 \).

Figure 13 depicts the transients of the errors \( \tilde{\theta}_i(t) \) for (3.5)—(a), (2.7)—(b), and (2.3)—(c).

The obtained transients confirmed the theoretical conclusions made in Remark 4. Indeed, under the conditions of the second statement of Proposition 2, the law (3.5), in contrast to (2.7) and (2.3), provided a monotonic decrease of the error \( \tilde{\theta}_i(t) \) over the corresponding time intervals when \( d_i(t) = 0 \):

\[
|\tilde{\theta}_1(t)| \leq \beta |\tilde{\theta}_1(0)|, \quad |\tilde{\theta}_2(t)| \leq \beta |\tilde{\theta}_2(1)|, \quad |\tilde{\theta}_3(t)| \leq \beta |\tilde{\theta}_3(2)|, \quad \beta \in (0; 1).
\]

Comparing the transients, the advantages of the law (3.5) is seen over (2.7) and (2.3). As for (2.3), the law (3.5) ensured the first-order type transients of \( \tilde{\theta}_i(t) \), \( \forall i \in \{1, 2, 3\} \), throughout the experiment. Compared to (2.7), the law (3.5) converged not just over the time ranges \([1; 1.165]\) and \([2; 2.14]\), but for all \( t \geq 0 \).

Figure 14 depicts the transients of \( \tilde{z}(t) \) for laws (3.5)—(a), (2.7)—(b), and (2.3)—(c).

The transients in Fig. 14 confirm the exponential recovery of the tracking error \( \tilde{z}(t) \) to its equilibrium point proved in Theorem 3 when \( \tilde{\theta}(t) \in s\text{-PE} \) and Assumption 3 was met.

Figure 15 shows transient curve of the \( \tilde{\Theta}(t) \) norm.

Figure 15 validates the exponential recovery of the parameter error \( \tilde{\Theta}(t) \) to its equilibrium point when \( \tilde{\theta}(t) \in s\text{-PE} \) and Assumption 3 was met, which followed the conclusions made in Theorem 3.

Since the number of rank switches was finite \( j \leq j_{max} = 4 < \infty \), according to the results of Corollary 5, the exponential recovery of the errors \( \tilde{z}(t) \) and \( \tilde{\Theta}(t) \) to their equilibrium points is equivalent to exponential stability.
Figure 16 presents transients of $\|\hat{\theta}(t)\|$ or the laws (3.5)—(a), (2.7)—(b), and (2.3)—(c).

The simulation results shown in Fig. 16 validate the conclusions made in Theorem 1. Indeed, when $\|\hat{\theta}(t)\| < \theta_{\text{max}}$, the law (3.5) was not convergent (when $\|\hat{\theta}(t)\| = \theta_{\text{max}}$, it was quasi-convergent), and the error norm $\|\hat{\theta}(t)\|$ could become greater than $\|\hat{\theta}(t^*)\|$.
Then it was set that \( \theta_0 = [0 \ 5 \ 0]^T \) which, according to Theorem 1, ensured that sufficient conditions of convergence of the law (3.5) were met:

\[
\beta_1 = \|\tilde{\theta}(t)\|\|\theta\|^{-1} \approx \frac{18}{15} = 1.2; \quad \frac{1}{\beta_1} + e^{-\gamma\delta} = \frac{1}{1.2} + e^{-5} = 0.84 \in (0; 1).
\]

Figure 17 shows the transients of \( \|\tilde{\theta}(t)\| \) obtained under such initial conditions, when the laws (3.5), (2.7), and (2.3) were applied.

The transient of \( \|\tilde{\theta}(t)\| \) for the law (3.5) confirmed the conclusions made in Theorem 1. The goal (2.2) was achieved when \( \bar{\psi}(t) \in s\text{-FE} \) and sufficient conditions were met, and \( \tilde{\theta}(t) \) did exponentially converge to the set with the bound \( \theta_{\max} \). Considering (2.7), such properties held only when \( \bar{\psi}(t) \in FE \).

Thus, the numerical experiments confirmed all theoretically stated properties of the proposed law (3.5). The results of Subsection 3.1 are valid in the general case \( \bar{\psi}(t) \in s\text{-FE} \), and the results of Subsections 3.2 and 3.3 are applicable under Assumptions 2 and 3, respectively.

5. CONCLUSIONS

In order to solve the identification problem of the unknown time-invariant parameters of a linear regression equation under the regressor semi-finite excitation, a procedure of dynamic regressor extension, regularization and mixing was proposed that generalized the well-known DREM method and extended the area of its applicability as far as practical scenarios were concerned.

In contrast to the conventional gradient-based identification law (2.3), the proposed procedure provided element-wise monotonicity of errors when Assumption 2 was met and exponential convergence of the tracking error of the function (2.1) when the regressor was semi-persistently exciting with the rank not less than one.

In contrast to the conventional DREM procedure, the developed one, firstly, relaxed the requirement of the regressor finite excitation previously required for convergence of (2.7) and ensured that the unknown parameters identification error decreased when the weaker condition of the regressor semi-finite excitation was met, and secondly, guaranteed exponential convergence of the regressand (2.1) tracking error when the regressor was semi-persistently exciting with the rank not less than one.

The scope of future research is the analysis and development of the dynamic regressor extension, regularization, and mixing procedure to solve the following problems:

—synthesis of the adaptive control schemes with relaxed requirements of the regressor excitation to ensure exponential convergence of the reference model tracking error;

—development of adaptive state observers with relaxed regressor excitation requirements for exponential convergence of plant states tracking error to zero;

—using partial identifiability conditions (see Proposition 2 and Figs. 13a, 7a, 2a) to identify the full vector of plant unknown parameters in case of over-parameterization;

—based on Proposition 2 and simulation results (see Figs. 13a, 7a, 2a), development of an identification law that does not require a finite or persistent excitation of the regressor to provide exponential convergence of the identification error of the full vector of unknown parameters.
**Proof of Proposition 1.** The lower bounds of the regressor $\omega(t)$ are written on the basis of Corollaries 1–4:

$$\bar{\Phi}(t) \in \text{PE} \iff \forall t \geq kT \quad \omega(t) = \det(\Phi(t)) = \prod_{i=1}^{n} \lambda_i(t) \geq \lambda_{\text{min}}(t) > \mu^n > 0,$$

$$\bar{\Phi}(t) \in \text{FE} \iff \forall t \in [t_0; t_0 + \delta] \subset [t_r^*; t_s^*] \quad \omega(t) = \prod_{i=1}^{n} \lambda_i(t) \geq \lambda_{\text{min}}(t) > \mu^n > 0,$$

$$\bar{\Phi}(t) \in \text{s-PE} \iff \forall t \geq kT \quad \omega(t) = e^{-\sum_{i=1}^{n} \lambda_i(t)} \geq \min\{\lambda_{\text{min}}(t), \varepsilon^n\} > 0,$$

$$\bar{\Phi}(t) \in \text{s-FE} \iff \forall t \in [t_0; t_0 + \delta] \subset [t_r^*; t_s^*] \quad \omega(t) = e^{-\sum_{i=1}^{n} \lambda_i(t)} \geq \min\{\lambda_{\text{min}}(t), \varepsilon^n\} > 0,$$

as was to be proved in Proposition 1.

**Proof of Theorem 1.** (1) As, following Corollaries 1 and 2, the following implications hold when $\bar{\Phi}(t) \in \text{FE}/\bar{\Phi}(t) \in \text{PE}$:

$$\bar{\Phi}(t) \in \text{PE} \iff \forall t \geq kT \quad \lambda_{\text{min}}(t) > \mu > 0,$$

$$\bar{\Phi}(t) \in \text{FE} \iff \forall t \in [t_0; t_0 + \delta] \subset [t_r^*; t_s^*] \quad \lambda_{\text{min}}(t) > \mu > 0,$$

then, when $\bar{\Phi}(t) \in \text{FE}/\bar{\Phi}(t) \in \text{PE}$, in accordance with (3.2), zero eigenvalues in $\Lambda(t)$ are not substituted.

$$\Xi(t) = \cdot \quad \text{for the regressor matrix} \quad \Phi(t),$$

then it holds for the unknown parameters $\Theta$ that $\Theta = \theta$ owing to $\bar{\Phi}(t)$, and the identification law (3.5) coincides with (2.7) up to the definition of the adaptive gain $\gamma$, from which it follows that (3.5) ensures $b_1–b_5$ when $\bar{\Phi}(t) \in \text{FE}/\bar{\Phi}(t) \in \text{PE}$.

(2) The following function, in which time arguments are omitted for the sake of brevity, is introduced:

$$\forall t \in [t_r^*; t_s^*] \quad L = \bar{\Theta}^T \bar{\Theta}.$$

Equation (A.2) is differentiated along the solutions of (3.5) to obtain:

$$L = -2\bar{\Theta}^T (\gamma_0 (\omega \bar{\Theta} - \omega \theta + \omega V \bar{\Lambda}^{-1} \Xi V^T \theta)) = -2\bar{\Theta}^T \gamma_0 \bar{\Theta} - 2\bar{\Theta}^T \gamma_0 \bar{\Theta} \bar{\Lambda}^{-1} \Xi V^T \theta.$$

Considering Assumption 1 and the definition of $\gamma$, the upper bound of (A.3) for all $t \in [t_0; t_0 + \delta] \subset [t_r^*; t_s^*]$ is written as:

$$\dot{L} \leq -2\bar{\Theta}^T \frac{\gamma_0}{\omega} \bar{\Theta} - 2\bar{\Theta}^T \frac{\gamma_0}{\omega} \bar{\Theta} V \bar{\Lambda}^{-1} \Xi V^T \theta \leq -2\bar{\Theta}^T \gamma_0 \bar{\Theta} - 2\bar{\Theta}^T \gamma_0 \bar{\Theta} \bar{\Lambda}^{-1} \Xi V^T \theta \leq -2\gamma_0 \|\bar{\Theta}\|^2 + 2\gamma_0 \|\theta\|_{\text{max}}.$$

Here, spectral norm of the multiplier $V \bar{\Lambda}^{-1} \Xi V^T$, which value is one as the matrices $V$ and $V^T$ are orthogonal ones, is calculated to obtain (A.4).

Assuming that $a = \sqrt{2\gamma_0 \|\bar{\Theta}\|}$, $b = \sqrt{2\gamma_0 \theta_{\text{max}}}$ and using the inequality $-a^2 + ab \leq -\frac{1}{2}a^2 + \frac{1}{2}b^2$, it is obtained from (A.4):

$$L \leq -\gamma_0 \|\bar{\Theta}\|^2 + \gamma_0 \theta_{\text{max}}^2.$$

The solution of the differential inequality (A.5) for all $t \in [t_0; t_0 + \delta]$ is written as:

$$\forall t \in [t_0; t_0 + \delta] \quad L \leq e^{-\gamma_0(t-t_0)} \|\bar{\Theta}(t_0)\|^2 + \theta_{\text{max}}.$$

Considering (A.6), $L = \|\bar{\Theta}\|^2$ and the fact that for all $c$, $d$ the inequalities $\sqrt{c^2 + d^2} \leq \sqrt{c^2} + \sqrt{d^2}$ hold, we obtain:

$$\|\bar{\Theta}(t_0 + \delta)\| \leq e^{-0.5\gamma_0} \|\bar{\Theta}(t_0)\| + \theta_{\text{max}}.$$
As for the most conservative case, it holds that \( \omega(t) \equiv 0 \) for all \( t \in \{ t^*_L, t^*_L + \delta; t_e \} \); therefore, the inequalities \( \| \tilde{\Theta}(t^*_L) \| > \| \tilde{\Theta}(t^*_L) \| \geq \| \tilde{\Theta}(t^*_L + \delta) \| \) also hold, using which (A.7) is rewritten as:
\[
\| \tilde{\Theta}(t_e) \| \leq e^{-0.5\gamma_0 \delta} \| \tilde{\Theta}(t^*_L) \| + \Theta_{\text{max}}. \tag{A.8}
\]

The premise (2.1) is substituted into (A.8) to obtain:
\[
\| \tilde{\Theta}(t_e) \| \leq e^{-0.5\gamma_0 \delta} \| \tilde{\Theta}(t^*_L) \| + \dfrac{1}{\beta_1} \| \tilde{\Theta}(t^*_L) \|. \tag{A.9}
\]

Hence, the choice of \( \gamma_0 \) on the basis of the condition
\[
0 < e^{-0.5\gamma_0 \delta} + \dfrac{1}{\beta_1} < 1 \iff \gamma_0 > \dfrac{2 \ln \left( 1 - \dfrac{1}{\beta_1} \right)}{\delta} \tag{A.10}
\]
allows one to ensure that the premise (2.2) also holds and, as a consequence, obtain the following:
\[
\| \tilde{\Theta}(t_e) \| \leq \left( e^{-0.5\gamma_0 \delta} + \dfrac{1}{\beta_1} \right) \| \tilde{\Theta}(t^*_L) \|. \tag{A.11}
\]

which means that the error \( \tilde{\Theta}(t) \) decreases over the time range \( [t^*_L; t_e] \).

The substitution of (A.11) into the upper bound of \( \bar{z}(t_e) \) yields:
\[
\| \bar{z}(t_e) \| \leq \bar{\varphi}_{\text{max}} \| \tilde{\Theta}(t_e) \| \leq \bar{\varphi}_{\text{max}} \left( e^{-0.5\gamma_0 \delta} + \dfrac{1}{\beta_1} \right) \| \tilde{\Theta}(t^*_L) \| = \beta \| \bar{z}(t^*_L) \|, \tag{A.12}
\]
which completes the proof of the second statement and verifies the convergence of (3.5) when \( \bar{\varphi}(t) \in s\text{-FE} \) and the premises (2.1) and (2.2) hold.

(3) The derivative of \( \Theta(t) \) is calculated to prove the third statement:
\[
\dot{\Theta}(t) = -\gamma(t) \omega(t) \dot{\Theta}(t) - \Theta(t). \tag{A.13}
\]

The general solution of the differential equation (A.13) is:
\[
\Theta(t) = \phi(t, t_0) \Theta(t_0) - \int_{t_0}^{t} \phi(t, \tau) \dot{\Theta}(\tau) d\tau, \tag{A.14}
\]
where \( \phi(t, s) = e^{-\gamma(t)(t_0, t) d\tau} \).

As, owing to \( \sqrt{\gamma(t) \in L_2, \frac{\sqrt{\gamma(t)}}{\omega(t) \in L_2,} \) and \( \omega(t) \in L_2, \) for all possible switches of the nonlinear operator in (3.5) it is true that \( \sqrt{\gamma(t) \in L_2,} \) then the function \( \phi(t, s) \) has the following properties:
\[
\sqrt{\gamma(t) \in L_2} \iff \begin{cases} 0 < \phi(t, s) \leq 1, \\ \lim_{t \to \infty} \phi(t, s) = 0. \end{cases} \tag{A.15}
\]

Using the first property, the upper bound of (A.14) is obtained:
\[
\Theta(t) \leq \phi(t, t_0) \Theta(t_0) - \Theta(t). \tag{A.16}
\]

On the basis of (A.16) and definitions \( \dot{\Theta}(t) = \dot{\theta}(t) + d(t) \), \( \Theta(t) = \theta - d(t) \) we have:
\[
\dot{\theta}(t) \leq \phi(t, t_0) \Theta(t_0) - \theta. \tag{A.17}
\]

From this, based on the second property of (A.15), it follows that \( \lim_{t \to \infty} \| \tilde{\Theta}(t) \| \leq \Theta_{\text{max}}, \) which completes the proof of the third statement of the theorem.
(4) When the condition \( \overline{q}(t) \in s\text{-PE} \) is met, in accordance with the third statement of Proposition 1 for all \( t \geq kT \) it holds that \( \omega(t) \geq \min(\lambda, \varepsilon) > 0 \) and, consequently, the function \( \phi(t, kT) \) is written as:

\[
\phi(t, kT) = e^{-\gamma_0(t-kT)}.
\]  

(A.18)

Then, having solved (A.13) for all \( t \geq kT \), the following is obtained in a similar manner to (A.14)–(A.17):

\[
\frac{1}{2} \| \hat{\Theta}(t) \|^2 \leq e^{-\gamma_0(t-kT)} \| \hat{\Theta}(kT) \|^2 + \theta_{\max},
\]  

(A.19)

from which it follows that, when \( \overline{q}(t) \in s\text{-PE} \), the errors \( \hat{\Theta}(t) \) exponentially convergence to the set with the bound \( \theta_{\max} \), which completes the proof of the theorem.

**Proof of Theorem 2.** (I) To prove the first statement of Theorem 2, Eq. (3.4) is written in the element-wise form:

\[
\Upsilon_i(t) = \omega(t) \Theta_i, \quad \forall i \in \{1, \ldots, n\}.
\]  

(A.20)

Given (A.20), the law (3.5) for all \( i \in \{1, \ldots, n\} \) is written as follows:

\[
\dot{\Theta}_i(t) = \hat{\Theta}_i(t) = -\gamma(t) \omega(t)(\omega(t) \dot{\hat{\Theta}}_i(t) - \omega(t) \Theta_i) = -\gamma(t) \omega(t) \Theta_i.
\]  

(A.21)

As \( \gamma(t) \omega(t) > 0 \), then \( \text{sgn} \{\dot{\hat{\Theta}}_i(t)\} = \text{const} \), and it holds for \( \hat{\Theta}_i(t) \) that \( |\hat{\Theta}_i(t)| \leq |\hat{\Theta}_i(t_0)| \), \( \forall t_a \geq t_b \), which was to be proved in part I of the theorem.

(II) When \( \overline{q}(t) \in s\text{-FE} \) and Assumption 2 is met, in accordance with Corollary 4 the solution of Eq. (A.13) over \( [t_b; t_a + \delta] \) is written as:

\[
\hat{\Theta}(t) = \phi(t, t_b) \hat{\Theta}(t_b) = e^{-\gamma_0(t-t_b)} \hat{\Theta}(t_b).
\]  

(A.22)

Considering the most conservative case, for all \( t \in \{t_b^*; t_b\}, \{t_b + \delta; t_e\} \) it holds that \( \omega(t) = 0 \); therefore, we have the inequalities \( \|\hat{\Theta}(t_b^*)\| \geq \|\hat{\Theta}(t_b)\|, \|\hat{\Theta}(t_e)\| \leq \|\hat{\Theta}(t_b + \delta)\| \), on the base of which the upper bound of \( \Theta(t) \) at the time instant \( t_e \) is obtained:

\[
\|\hat{\Theta}(t_e)\| \leq e^{-\gamma_0\delta} \|\hat{\Theta}(t_b^*)\|.
\]  

(A.23)

The definition \( \beta = e^{-\gamma_0\delta} \in (0; 1) \) is introduced into (A.23) to complete the proof that the error \( \hat{\Theta}(t) \) decreases over \( [t_b^*; t_e] \).

To prove the error \( \hat{z}(t) \) reduction, the correctness of the following implication owing to \( V_{1T}(t)V_2 = 0_{r \times r} \) is taken into consideration:

\[
y(t) = \Phi(t) \Theta = V(t) \Lambda(t)V_{1T}(t)(\Theta - V_2 V_2^T \Theta) = \Phi(t)(\Theta - V_2 V_2^T \Theta) = \Phi(t) \Theta
\]

\[
= \int_{t_b}^{t} e^{-\gamma_0(t-\tau)} \Phi(t) \Phi(t)^T \Theta \tau \tau d\tau
\]

\[
= \int_{t_b}^{t} e^{-\gamma_0(t-\tau)} \Phi(t) \Phi(t)^T \Theta \tau \tau d\tau
\]

\[
= \int_{t_b}^{t} e^{-\gamma_0(t-\tau)} \Phi(t) \Phi(t)^T \Theta \tau \tau d\tau
\]

\[
= \int_{t_b}^{t} e^{-\gamma_0(t-\tau)} \Phi(t) \Phi(t)^T \Theta \tau \tau d\tau
\]

\[
\Rightarrow \hat{z}(t) = \Phi(t) \Theta = \Phi(t)(\Theta - V_2 V_2^T \Theta) = \Phi(t) \Theta.
\]  

(A.24)

Then, considering (A.22), the upper bound of the tracking error is written as:

\[
\forall t \in [t_b; t_b + \delta] \|\hat{z}(t)\| \leq \bar{q}(t_b) e^{-\gamma_0(t-t_b)} \|\hat{\Theta}(t_b)\|.
\]  

(A.25)

from which, owing to (A.23), we immediately have:

\[
\|\hat{z}(t_e)\| \leq \bar{q}_{\max} \beta \|\hat{\Theta}(t_e)\| = \beta \|\hat{z}(t_e)\|
\]  

(A.26)

which was to be proved in part II.
(III) When Assumption 2 is met, for all \( t \in [\gamma_0, \infty) \) the solution of the error (A.13) is written as:

\[
\hat{\Theta}(t) = \phi(t, t_0) \Theta(t_0),
\]

from which, according to the second property of (A.15), it follows that:

\[
\sqrt{\gamma(t)} \alpha(t) \in L_2 \iff \lim_{t \to \infty} \|\hat{\Theta}(t)\| = 0,
\]

which holds for all possible variants of switches of the nonlinear operator (3.5) owing to \( \sqrt{\gamma_0} \notin L_2 \), \( \sqrt{\gamma_0} \notin L_2 \), and \( \alpha(t) \notin L_2 \).

Having applied the implication (A.28) to obtain the upper bound of (A.24), we have:

\[
\sqrt{\gamma(t)} \alpha(t) \in L_2 \iff \lim_{t \to \infty} \|\hat{\Theta}(t)\| = 0.
\]

Thus, all statements of the third part of Theorem 2 are proved.

(IV) When \( \Omega(t) \in s\text{-PE} \), then (A.18) holds \( \forall t \geq kT \), and therefore the following bound is obtained on the basis of (A.22):

\[
\forall t \geq kT, \quad \|\hat{\Theta}(t)\| \leq e^{-\gamma(t-kT)} \|\hat{\Theta}(kT)\|,
\]

which proves the exponential convergence of the error \( \hat{\Theta}(t) \) to zero for all \( t \geq kT \).

Having (A.30) at hand, considering the boundedness of \( \|\hat{\Theta}(t)\| \leq \bar{\Theta}_{\text{max}} \) and using (A.24), the exponential convergence of the error \( \hat{\Theta}(t) \) for all \( t \geq kT \) can be proved in the similar way to (A.25), which completes the proof of Theorem 2.

**Proof of Theorem 3.** When \( \Omega(t) \in s\text{-PE} \), on the basis of the third statement of proved Proposition 1 for all \( t \geq kT \) \( \alpha(t) \geq \min(\lambda_{\text{min}}^n(t), \epsilon^n) > 0 \) holds, and therefore Eq. (A.13) is written as:

\[
\forall t \geq kT, \quad \hat{\Theta}(t) = -\gamma_0 \hat{\Theta}(t) - \hat{\Theta}(t).
\]

Owing to Assumption 3, the derivative \( \hat{\Theta}(t) \) is written as follows according to (3.7):

\[
\hat{\Theta}(t) = \sum_{j=1}^{\infty} \bar{\Delta}_j \hat{\Theta}(t - t_j).
\]

Considering (A.32), the solution of the differential Eq. (A.31) is obtained:

\[
\forall t \geq kT, \quad \hat{\Theta}(t) = e^{-\gamma(t-kT)} \Theta(kT) - \int_{kT}^{t} e^{-\gamma(t-\tau)} \sum_{j=1}^{m} \bar{\Delta}_j \hat{\Theta}(\tau - t_j) d\tau.
\]

Following the sifting property of the Dirac function, for any differentiable function \( f(t) \) we have:

\[
\int_{t_0}^{t} f(\tau) \delta(\tau - t_j) d\tau = f(t_0) h(\tau - t_j) \bigg|_{t_0}^{t} = f(t_0) h(t - t_j) - f(t_j) h(t - t_j) = f(t_j) h(t) - f(t_j) h(t - t_j).
\]

On the basis of (A.34) Eq. (A.33) is rewritten as:

\[
\forall t \geq kT, \quad \hat{\Theta}(t) = e^{-\gamma(t-kT)} \Theta(kT) - \sum_{j=1}^{m} e^{-\gamma(t-t_j)} \bar{\Delta}_j h(t - t_j).
\]

Having multiplied (A.35) by \( \hat{\Theta}^T(kT) \), it is obtained:

\[
\forall t \geq kT, \quad \hat{\Theta}^T(kT) \hat{\Theta}(t) = e^{-\gamma(t-kT)} \|\Theta(kT)\|^2 - \sum_{j=1}^{m} e^{-\gamma(t-t_j)} \bar{\Delta}_j h(t - t_j).
\]
The term $e^{-\gamma(t-kT)}\|\tilde{\Theta}(t)\|^2$ is put outside the brackets in the right-hand side of Eq. (A.36) to obtain for all $t \geq kT$ that:

$$\tilde{\Theta}(t) = \left(1 - \frac{1}{\|\tilde{\Theta}(kT)\|^2} \sum_{j=1}^{\infty} e^{-\gamma(kT-t)}\tilde{\Theta}^T(kT)\Delta_j h(t-t_j)\right)e^{-\gamma(t-kT)}\tilde{\Theta}(kT),$$

(A.37)

where $\|\tilde{\Theta}(kT)\| \neq 0$ since for all $t \in [0; kT]$ $\alpha(t) \equiv 0 \Rightarrow \tilde{\Theta}(t) = 0 \Rightarrow \|\tilde{\Theta}(kT)\| \geq \|\tilde{\Theta}(0)\|$. Equation (A.37) allows one to have the first expression from (3.8) up to the following notation:

$$a(t_j) = \left|1 - \frac{1}{\|\tilde{\Theta}(kT)\|^2} \sum_{j=1}^{\infty} e^{-\gamma(kT-t)}\tilde{\Theta}^T(kT)\Delta_j h(t-t_j)\right|.$$  

(A.38)

So the exponential recovery of the parameter error $\tilde{\Theta}(t)$ to its equilibrium point is proved.

Having (A.24) at hand, the upper bound of the tracking error $|\bar{z}(t)|$ is written as:

$$\forall t \geq kT \quad |\bar{z}(t)| \leq a(t_j)\mathcal{P}_{\max} e^{-\gamma(t-kT)}\|\tilde{\Theta}(kT)\| = a(t_j)e^{-\gamma(t-kT)}|\bar{z}(kT)|.$$

(A.39)

Therefore, the exponential recovery of the error $\bar{z}(t)$ to its equilibrium point is also proved.

If, additionally, for $a(t_j)$ there exists an upper bound $a_{\max}$, then it is immediately obtained from (3.8) that:

$$\lim_{t \to \infty} \|\tilde{\Theta}(t)\| \leq \lim_{t \to \infty} a_{\max} e^{-\gamma(t-kT)}\|\tilde{\Theta}(kT)\| = 0$$

$$\lim_{t \to \infty} |\bar{z}(t)| \leq \lim_{t \to \infty} \left(a_{\max} e^{-\gamma(t-kT)}\|\tilde{\Theta}(kT)\|\right) = \lim_{t \to \infty} \left(a_{\max} e^{-\gamma(t-kT)}|\bar{z}(kT)|\right) = 0.$$  

(A.40)

Hence, the tracking error $\bar{z}(t)$ and the parameter error $\tilde{\Theta}(t)$ are exponentially stable, which completes the proof of Theorem 3.

**Proof of Corollary 5.** According to the first statement of Corollary 5, it is assumed that the number of $\Theta(t)$ changes is finite: $j \leq j_{\max} < \infty$.

Then the following upper bound of the function $a(t_j)$ is obtained:

$$a(t_j) = \left|1 - \frac{1}{\|\tilde{\Theta}(kT)\|^2} \sum_{j=1}^{\infty} e^{-\gamma(kT-t)}\tilde{\Theta}^T(kT)\Delta_j h(t-t_j)\right| \leq 1 + \frac{1}{\|\tilde{\Theta}(kT)\|^2} \sum_{j=1}^{\infty} e^{-\gamma(kT-t)}\tilde{\Theta}^T(kT)\Delta_j h(t-t_j)$$

$$\leq 1 + \frac{1}{\|\tilde{\Theta}(kT)\|^2} \sum_{j=1}^{\infty} \|\Delta_j\| e^{-\gamma(kT-t)}h(t-t_j).$$  

(A.41)

As, when $j$ is finite, the number of time instants $t_j$ is also finite, then the exponential multiplier in the sum (A.41) is bounded, and the following definition holds:

$$a(t_j) \leq 1 + \frac{1}{\|\tilde{\Theta}(kT)\|^2} \sum_{j=1}^{\infty} \|\Delta_j\| e^{-\gamma(kT-t)}h(t-t_j) = a_{\max},$$

(A.42)

which was to be proved in the first part of the corollary.

To prove the second statement of the corollary, the upper bound of $\|\Delta_j\|$ is taken into consideration, and the upper bound of $a(t_j)$ is obtained similarly to (A.42), but under the condition of the infinite number of switches:

$$a(t_j) \leq 1 + \frac{1}{\|\tilde{\Theta}(kT)\|^2} \sum_{j=1}^{\infty} e^{-\gamma(kT-t)}\tilde{\Theta}^T(kT)\Delta_j h(t-t_j) \leq 1 + \sum_{j=1}^{\infty} c(t_j)h(t-t_j).$$

(A.43)
The series from (A.43) is of positive terms, and all its subsums are bounded because of monotonicity $0 < c(t_{j+1}) \leq c(t_j)$, and, therefore, $1 + \sum_{j=1}^{\infty} c(t_j) h(t - t_j) \leq a_{\text{max}}$, which completes the proof of Corollary 5.

**Proof of Proposition 2.** As, when $(t)$ FE/$(t)$ PE, the following implications hold according to Corollaries 1 and 2:

$$
\bar{\varphi}(t) \in \text{PE} \iff \forall t \geq kT \quad \lambda_{\text{min}}(t) > \mu > 0,
$$

$$
\bar{\varphi}(t) \in \text{FE} \iff \forall t \in [t_0, t_0 + \delta] \subset [t_j, t_{j+1}] \quad \lambda_{\text{min}}(t) > \mu > 0,
$$

then, when $\bar{\varphi} = 0$, according to (3.3) we have $\Xi(t) = 0$, as a result $\bar{\Xi}(t) = 0$, and, consequently, $\bar{\varphi}(t) \in \text{FF/PE} \Rightarrow d(t) = 0 \Rightarrow \Theta(t) = 0$, which completes the proof of statement (a) of Proposition 2.

The necessity of conditions $(t) \in \text{s-FE/PE}$ follows from the fact that only if $0 < r < n$, the premises of the statement (b) are consistent ($\exists p > 0 \sum_{i=1}^{n-p} w_i \varphi_i(t) = 0_n, w_i \neq 0$). The necessity of the condition $n > 2$ follows from the contradiction, which occurs when $n = 2$ in general case ($\varphi_1(t) \neq 0_n$):

$$
w_i \varphi_i(t) + w_j \varphi_j(t) = 0_n, \quad w_i \neq 0, \quad w_j = 0.
$$

The next step is to prove the necessity and sufficiency of the following condition to ensure that $\exists M \subset \{1, \ldots, n\}, |M| = p, \forall i \in M, \Theta_i = 0$:

$$
\sum_{i=1}^{n-p} w_i \varphi_i(t) + \sum_{j=n-p+1}^{n} w_j \varphi_j(t) = 0_n, \quad w_i \neq 0, \quad w_j = 0.
$$

(A.44)

**Necessity.** To begin with, it should be noted that according to (3.5), the elements of the vector of new unknown parameters $\Theta$ coincide with the elements of the vector of original parameters $\theta$ if the corresponding elements of the vector $d$ are equal to zero. Therefore, $d$ is considered in more detail. If $\bar{\varphi} > 0$, the multiplication $\bar{\Xi}(t)$ has the following structure:

$$
\bar{\Xi}(t) = \begin{bmatrix} \Lambda_{11}(t) & 0_{r \times p} & 0_{p \times 1} \\ 0_{p \times r} & 0_{r \times r} & I_r \\ 0_{p \times r} & e^{-1} I_r & 0_{r \times r} \end{bmatrix} = \begin{bmatrix} 0_{r \times r} & 0_{r \times r} \\ 0_{p \times r} & e^{-1} I_r \\ 0_{p \times r} & I_r \end{bmatrix}
$$

(A.45)

Then, owing to the notation (3.4), the definition of $d$ is rewritten as:

$$
d = V(t) \bar{\Xi}^{-1}(t) \Xi(t) V^T(t) \Theta = V_2 V_2^T \Theta = [d_1, \ldots, d_n]^T,
$$

(A.46)

from which it follows that $d$ has $p$ zero elements and columns of the matrix $V_2 V_2^T$ is $p$, which, in turn, is satisfied when the matrix $V_2$ has $p$ zero rows.

Following the definition of the singular decomposition of a positively semi-definite symmetric matrix [15, 16], the matrix $V_2$ can be obtained as a solution of a homogeneous system of linear algebraic equations:

$$
\varphi(t) V^k_2 = \sum_{i=1}^{n} \varphi_i(t) V^k_2 = 0_n, \quad \forall k \in \{1, \bar{\varphi}\},
$$

(A.47)

where $V^k_2$ is the $k$th column of the matrix $V_2$.

To prove the necessity of the condition (A.44), it is to be shown that if $w_j \neq 0$, then the vector $V^k_2$, $\forall k \in \{1, \bar{\varphi}\}$, does not contain zero elements.

The expression (A.47) can be rewritten in the following equivalent form (taking into account the ortho-normality of $V^k_2$, $\forall k \in \{1, \bar{\varphi}\}$):

$$
\varphi(t) V^k_2 = \sum_{i=1}^{n} \varphi_i(t) V^k_2 = \frac{1}{\sum_{i=1}^{n} w_i^2} \sum_{i=1}^{n} w_i \varphi_i(t) = \frac{1}{\sum_{i=1}^{n} w_i^2} \left( \sum_{i=1}^{n-p} w_i \varphi_i(t) + \sum_{j=n-p+1}^{n} w_j \varphi_j(t) \right)
$$

(A.48)

$$
\sum_{i=1}^{n-p} \varphi_i(t) + \sum_{j=n-p+1}^{n} \varphi_j(t) = 0_n.
$$

(A.49)
Since we consider only nontrivial solutions to find $V^k_2$, if the condition (A.44) is not satisfied, the set of solutions is given as follows:

\[
V^k_2 = \frac{w_k}{\sqrt{\sum_{i=1}^{n} w_i^2}} \neq 0; \quad V^k_2 = \frac{w_j}{\sqrt{\sum_{i=1}^{n} w_i^2}} \neq 0,
\]

and then $V^k_2$, $\forall k \in \{1, \ldots, m\}$, does not include zero elements and, consequently, $\exists d_j = 0 \Rightarrow \exists M \subset \{1, \ldots, n\}$, $|M| = p$, $\forall i \in M$, $\Theta_j = \Theta_i$, which completes the proof of necessity of the condition (A.45).

**Sufficiency.** Following the statement of the proposition, when the condition (A.44) is met, the solution set of the equation of the form (A.47) is defined as follows:

\[
V^k_2 = \frac{w_k}{\sqrt{\sum_{i=1}^{n} w_i^2}} \neq 0; \quad V^k_2 = \frac{w_j}{\sqrt{\sum_{i=1}^{n} w_i^2}} = 0,
\]

and then the vector $V^k_2$, $\forall k \in \{1, \ldots, m\}$, includes $p$ zero elements and, consequently, $\exists M \subset \{1, \ldots, n\}$, $|M| = p$, $\forall i \in M$, $\Theta_j = \Theta_i$, which completes the proof of sufficiency of the condition (A.44).

Thus, the condition (A.44) is necessary and sufficient for the identifiability of $p$ elements of the unknown parameters vector $\theta$, which completes the proof of the second statement of Proposition 2.

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