Using Biased Coins as Oracles

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While it is well known that a Turing machine equipped with the ability to flip a fair coin cannot compute more than a standard Turing machine, we show that this is not true for a biased coin. Indeed, any oracle set $X$ may be coded as a probability $p_X$ such that if a Turing machine is given a coin which lands heads with probability $p_X$ it can compute any function recursive in $X$ with arbitrarily high probability. We also show how the assumption of a non-recursive bias can be weakened by using a sequence of increasingly accurate recursive biases or by choosing the bias at random from a distribution with a non-recursive mean. We conclude by briefly mentioning some implications regarding the physical realisability of such methods.

Keywords: Probabilistic Turing machine, hypercomputation, oracle, qubit.

1 INTRODUCTION AND MOTIVATION

The Turing machine is well known to be a very robust model of computation. In almost all textbooks on the theory of computation, one can find a list of extensions to the Turing machine that offer it more primitive resources, such as extra tapes or nondeterminism, and yet do not give it the ability to compute any additional functions. Amongst such resources it is not uncommon to find references to probabilistic methods such as coin tossing.

These methods can be made precise with the introduction of the probabilistic Turing machine or PTM [3]. A PTM is a standard Turing machine with a special randomising state. When the machine is in this state, the transition to a new state is not governed by what is on the tape, but by a random event.
A fair coin is tossed and the machine goes to the specified 1-state if the coin comes up heads and the 0-state if it comes up tails.

Unlike a Turing machine, a PTM will not necessarily return the same output when run multiple times on the same input. Care must therefore be taken in defining what it means for a function to be computed by a PTM. One way is to say that a PTM computes a given function, \( f \), if when given \( x \) as input, along with a measure of accuracy \( j \in \mathbb{N} \), it produces \( f(x) \) with probability at least \( 1 - \frac{1}{2^j} \). By this definition, a function is computable by a PTM if and only if it can be computed with arbitrarily high confidence. Alternatively, we could relax this definition and say that a PTM computes \( f \) if and only if when given \( x \) as input, it produces \( f(x) \) with some probability greater than \( \frac{1}{2} \).

It is quite easy to see that with either definition, a PTM computes only the recursive functions. For any PTM \( P \), there is a Turing machine \( T \) that simulates it. \( T \) simulates each branch of the computation in parallel and keeps track of their respective probabilities. \( T \) also keeps a table which associates outputs with their probabilities. When a branch halts and returns some value \( y \), \( T \) creates a new position in the table for \( y \) and stores the probability of that branch occurring. If a branch has already halted with output \( y \), \( T \) simply adds the new probability of producing \( y \) to the old value. After each update to the table, \( T \) checks whether the new value for \( y \) is greater than \( P \)'s threshold (\( \frac{1}{2} \) or \( 1 - \frac{1}{2^j} \)) and halts returning \( y \) if this is so. In this way, \( T \) halts with output \( y \) if and only if \( P \) returns \( y \) with sufficient probability.

This argument can also be extended to deal with more complicated probabilistic methods. For example, we could allow biased coins where the chance that heads comes up is some given rational number. We could even allow the bias to be any recursive real number (as defined in Section 2 of this paper). In each case, \( T \) can still keep track of the probability of each computation branch and test to see whether an output occurs with high enough probability to be deemed the output of the PTM.

It is important to ask, however, what can be computed if non-recursive probabilities are used. In this paper, we show that allowing coins with non-recursive biases makes the above argument fail quite spectacularly. We first show that a PTM can compute arbitrarily accurate estimates to the bias on its coin and then strengthen this to computing arbitrarily many bits of the binary expansion of the bias. From this, we reach several strong theorems about the power of PTMs, showing in particular that there is a single PTM that acts as a universal \( o \)-machine: when equipped with a probability coding

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1 Since this paper was written, an article by Santos [8] has been brought to our attention wherein a result similar to our Theorem 4.1 was obtained. However, it is our opinion that Santos’s proof is incomplete, lacking an explanation of how the binary expansion of the probability can be computed from the rational approximations. In any event, we think the present account is useful for its further results, discussion of the physical implications, and clarity of presentation.
a given oracle, it simulates a given \( o \)-machine with that oracle on a given input to a given level of confidence. Thus, the addition of randomness to the resources of a Turing machine expands its set of computable functions. Only when the coins are restricted to recursive biases do they offer no additional power.

In the remaining sections, we show two ways in which the same results are possible with slightly weakened resources. Specifically, we show how a sequence of rationally biased coins can be used, so long as the biases converge effectively to a non-recursive real or the biases are drawn at random from a distribution with a non-recursive mean. Finally, we point to some interesting physical applications in which these types of probabilistic methods seem to be consistent with quantum mechanics.

2 APPROXIMATING \( p \) TO ARBITRARY ACCURACY

The natural way to approximate the probability, \( p \), that the coin will land heads, is to look at the average number of heads in \( n \) tosses. By the weak law of large numbers, this value (which we will denote by \( \hat{p} \)) approaches \( p \) as \( n \) approaches infinity. However, to approximate \( p \) effectively, we need to know how fast this convergence is likely to be. This can be expressed by asking how many tosses are required before \( \hat{p} \) is within a given distance of \( p \) with a given level of confidence. Specifically, we will ask for a method of calculating \( n \) such that when at least \( n \) tosses are made, \( |\hat{p} - p| < \frac{1}{2^k} \) with probability at least \( 1 \) - \( \frac{1}{2^j} \) for given \( j, k \in \mathbb{N} \).

The probability distribution of possible values of \( \hat{p} \) for a given value of \( n \) is a binomial distribution with mean \( p \). The variance of \( \hat{p} \) is given by

\[
\sigma^2 = \frac{p(1-p)}{n} \tag{2.1}
\]

This variance depends upon the unknown value of \( p \), however since it has a maximum where \( p = \frac{1}{2} \), we can see that

\[
\sigma^2 \leq \frac{1}{4n} \tag{2.2}
\]

With this upper bound for the variance, we can use the Chebyshev inequality

\[
\forall \epsilon \geq 0 \quad P(|x - \mu| \geq \epsilon) \leq \frac{\sigma^2}{\epsilon^2} \tag{2.3}
\]

to form an upper bound for the probability of error

\[
\forall k \quad P \left( |\hat{p} - p| \geq \frac{1}{2^k} \right) \leq \frac{2^{2k}}{4n} \tag{2.4}
\]
Therefore, if we insist on a chance of error of at most \( \frac{1}{2^j} \), this can be achieved so long as

\[
\frac{2^k}{4n} \leq \frac{1}{2^j}
\]

(2.5)

\[
n \geq 2^{j+2k-2}
\]

(2.6)

Thus, for each value of \( p \in [0, 1] \) we can compute an approximation of \( p \) that is within an arbitrarily small distance of the true value with an arbitrarily high probability. More formally,

**Theorem 2.1.** There is a specific PTM that, when equipped with a probability \( p \), takes inputs \( j, k \in \mathbb{N} \) and outputs a rational approximation to \( p \) that is within \( \frac{1}{2^k} \) of the true value with probability at least \( 1 - \frac{1}{2^j} \).

**Proof.** The PTM simply tosses its coin \( 2^{j+2k-2} \) times and returns the ratio of heads to tails. By the argument above, this approximation will suffice. \( \square \)

This method of approximating a real number by successively accurate rational approximations can also be used to define what it means for a real to be computable by a (deterministic) Turing machine. For convenience, we say

**Definition 2.2.** \( \{x_n\} \) converges quickly to \( x \) if and only if \( |x_n - x| < \frac{1}{2^n} \) for all \( n \).

We can then define a recursive real:

**Definition 2.3.** \( x \in \mathbb{R} \) is recursive if and only if there is a Turing machine that takes \( n \in \mathbb{N} \) as input and returns \( x_n \in \mathbb{Q} \), where \( \{x_n\} \) converges quickly to \( x \).

The recursive reals given by this definition are well studied and include a great many of the reals actually encountered in mathematics, including all the algebraic numbers as well as \( \pi \) and \( e \). However, since there are uncountably many reals but only countably many Turing machines, it is clear that most of them are not recursive. If a PTM is equipped with one of these non-recursive reals as its probability, then our algorithm above shows that in a certain sense, this PTM can compute this real—a feat that is impossible with a deterministic Turing machine.

However, there is still some room to question whether the PTM of Theorem 2.1 actually computes its probability. Consider, for example, the following alternative definition of a recursive real.

**Definition 2.4.** \( x \in \mathbb{R} \) is recursive if and only if there is a Turing machine that takes no input and outputs a sequence \( \{x_n\} \) which converges quickly to \( x \).

This definition is evidently equivalent to the previous one when it comes to deterministic Turing machines, but it is not immediately clear that the equivalence holds for PTMs. While the PTM of Theorem 2.1 can compute each approximation to \( x \) with arbitrary accuracy, it is not clear that a PTM could output an infinite sequence of approximations with them all being correct with
arbitrarily high probability. However, we now show that this can be achieved by requiring each successive event to be more and more probable.

For a given minimum probability \( q \) that an entire infinite sequence of events will occur, we can set the probability of the \( i \)-th event occurring to \( q_i = q^{2^{-i}} \). It follows that the chance of all events occurring is

\[
\prod_{i=1}^{\infty} q_i = q^{\sum_{i=1}^{\infty} 2^{-i}} = q
\]  
(2.7)

In addition, we can consider the chance that all events in an infinite suffix of the sequence occur. The chance of all events after event \( N \) occurring is

\[
\prod_{i=N+1}^{\infty} q_i = q^{\sum_{i=N+1}^{\infty} 2^{-i}} = q^{2^{-N}}
\]  
(2.8)

Thus, for each \( \epsilon > 0 \), there is a value of \( N \) such that the probability of all events after event \( N \) occurring is within \( \epsilon \) of 1. Moreover, if we are just interested in there being some such \( N \) after which all events occur, this will happen with probability 1.

This construction can be applied in the case of our approximations to \( p \). In particular, we can find a new value \( j' \) as a function of \( j \) and \( k \) which can then be substituted into our formula for the number of required coin tosses.

\[
1 - \frac{1}{2^j} = \left( 1 - \frac{1}{2^j} \right)^{2^{-k}}
\]  
(2.9)

Using a Taylor expansion, we can see that for \( 0 \leq x, y < 1 \)

\[
(1 - x)^{y} < 1 - xy
\]  
(2.10)

and thus

\[
1 - \frac{1}{2^j} < 1 - \left( \frac{1}{2^j} \right) \left( \frac{1}{2^k} \right)
\]  
(2.11)

\[
1 - \frac{1}{2^j} < 1 - \left( \frac{1}{2^{j+k}} \right)
\]  
(2.12)

\[
j' < j + k
\]  
(2.13)

Putting this all together:

\textbf{Theorem 2.5.} There is a specific PTM that, when equipped with a probability \( p \), takes input \( j \in \mathbb{N} \) and outputs a sequence \( \{ \hat{p}_k \} \) that converges quickly to \( p \) with probability \( 1 - \frac{1}{2^j} \). Furthermore, with probability 1, there is some \( N \) such that \( |\hat{p}_k - p| < \frac{1}{2^k} \) for all \( k > N \).

\textit{Proof.} For each value in the sequence, the PTM simply tosses its coin \( 2^{j+3k-2} \) times and returns the ratio of heads to tails. By the argument above, these approximations will suffice. \( \square \)
3 COMPUTING THE BINARY EXPANSION OF $p$

The definitions of the previous section are not the only ways that the recursive reals can be defined. Instead of using converging sequences of rationals, we can use the original technique due to Turing [9] of using the base $b$ expansion. For simplicity, we use the binary expansion and only consider those reals in the unit interval.

Definition 3.1. $x \in \mathbb{R}$ is recursive if and only if there is a Turing machine that takes $n \in \mathbb{N}$ as input and returns $b_n$, the $n$-th bit of the binary expansion of $x$.

As before, we can rephrase this to speak of Turing machines that take no input:

Definition 3.2. $x \in \mathbb{R}$ is recursive if and only if there is a Turing machine that takes no input and returns the sequence $\{b_n\}$, corresponding to the binary expansion of $x$.

Both definitions run into an ambiguity in the case of dyadic rationals: those that can be expressed in the form $\frac{n}{2^m}$. For such numbers, there are two binary expansions so we adopt the convention of using the one containing an infinite number of 0s.

By extending our method for approximating $p$, we can also approximate the binary expansion of $p$. Unfortunately this will not be possible if $p$ is a dyadic rational, so for now consider the case where it is not, and $p$ thus has a unique infinite binary expansion in which both 0 and 1 occur infinitely many times. To compute the binary expansion of $p$, we need a method that takes inputs $j, l$ and gives us the value of $b_l$ with probability $1 - \frac{1}{2^k}$.

It may seem as though this can be achieved simply by computing $\hat{p}_{l+1}$ and taking its $l$-th bit, but problems arise when a run of consecutive 0s or 1s occurs around this point in the expansion. For instance, if we want the third bit and $\hat{p}_4 = 0.0111111$, then the true value of $p$ could be as low as 0.01101111 or as high as 0.10001111 and we can thus be certain of none of the bits. By using the following algorithm, which we shall call $A$, we can overcome this problem.

- $k := l$
- repeat
  - $k := k + 1$
  - compute $\hat{p}_k$ (by tossing the coin $2^{j+3k-2}$ times)
  - if $\hat{p}_k < 1$ and there are both a 0 and a 1 between the $l$-th and $k$-th bits of the expansion of $\hat{p}_k$ then output the $l$-th bit and halt

An analysis of $A$ is made somewhat complex by the fact that it involves random events and does not always give the correct output, but for now we...
will just consider the most probable case where the probabilistically generated sequence \( \{ \hat{p}_k \} \) converges quickly to \( p \). We can see that there must be a value of \( k \) for which \( \hat{p}_k \) is less than one and has both a 0 and a 1 between its \( l \)-th and \( k \)-th bits, for if there were not then \( \hat{p}_k \) would either be approaching a dyadic rational or failing to converge—each of which would contradict our assumptions. Therefore, so long as \( p \) is not a dyadic rational and \( \{ \hat{p}_k \} \) converges quickly to \( p \), \( A \) will always halt. When it does, the value of \( \hat{p}_k \) will be in the form

\[
\hat{p}_k = \cdot b_1 \ldots b_l 1 \cdots 10b_k \cdots \tag{3.1}
\]

or

\[
\hat{p}_k = \cdot b_1 \ldots b_l 0 \cdots 01b_k \cdots \tag{3.2}
\]

In either case, adding or subtracting a value smaller than \( \frac{1}{2^j} \) will not change any of the first \( l \) bits of \( \hat{p}_k \) and since \( p \) is within \( \frac{1}{2^j} \) of \( \hat{p}_k \), their first \( l \) bits must be identical.

It is important to note, however, that while all runs of 1s or 0s within the expansion of \( p \) must come to an end, they can be arbitrarily long, so the running time of \( A \) depends upon the value of \( p \). If the \( l \)-th bit of the expansion of \( p \) is followed by a run of \( m \) identical bits, then we must compute \( l + m \) values of \( \hat{p}_k \), requiring at most \( 2j + 3l + 3m - 1 \) coin tosses.

What about those cases where \( \{ \hat{p}_k \} \) does not converge quickly to \( p \)? This can be for two different reasons—either it converges to \( p \), but not as quickly as required or it does not converge to \( p \) at all. The first of these cases occurs with probability \( \frac{1}{2^j} \) and while it cannot cause \( A \) to fail to halt, it may well cause an incorrect output. The second case occurs only with probability 0, and may either cause an incorrect output or non-termination.

**Theorem 3.3.** There is a PTM that implements \( A \). Equipped with any non-dyadic probability \( p \), it takes positive integers \( j \) and \( l \), outputting the \( l \)-th bit of the binary expansion of \( p \) with probability greater than \( 1 - \frac{1}{2^j} \). The probability that it returns an incorrect answer is less than \( \frac{1}{2^j} \), while the probability that it does not terminate is 0.

**Proof.** Immediate. \( \square \)

We can also modify \( A \) to form \( A_\infty \) which takes only \( j \) as input and outputs the entire expansion of \( p \). In this case it outputs the \( l \)-th digit when it has output all prior digits and has found a value of \( \hat{p}_k \) with a 0 and a 1 between its \( l \)-th and \( k \)-th digits. \( A_\infty \) uses the high likelihood of \( \{ \hat{p}_k \} \) converging quickly to \( p \) to greater effect than \( A \), by generating the entire expansion with arbitrarily high probability.

**Theorem 3.4.** There is a PTM that implements \( A_\infty \). Equipped with any non-dyadic probability \( p \), it takes a positive integer \( j \), outputting the entire binary
expansion of $p$ with probability greater than $1 - \frac{1}{2^j}$. The probability that it outputs finitely many incorrect bits is less than $\frac{1}{2^j}$, while the probability that it outputs infinitely many incorrect bits or outputs only a finite number of bits is 0.

Proof. Immediate. □

4 USING THE BINARY EXPANSION OF $p$ AS AN ORACLE

In 1939, Alan Turing [10] introduced a very influential extension to his theoretical computing machines. Turing’s $o$-machines are standard Turing machines combined with a special ‘oracle’, which can answer questions about a particular set of natural numbers, called its oracle set. Like a PTM, an $o$-machine has a special query state and two answer states, but instead of the answer being given randomly, it corresponds to whether a certain number is in the oracle set. To specify the number whose membership is being questioned, a special symbol $\psi$ is inscribed twice on the tape and the number of squares between each inscription of $\psi$ is taken as the query to the oracle. Depending on which oracle set is given, an $o$-machine can compute different classes of functions, and they thus give rise to a notion of relative computability.

Corresponding to an $o$-machine with oracle $X$ we can construct a PTM with probability $p_X$ where the $n$-th digit of the binary expansion of $p_X$ is 1 if $n \in X$ and 0 otherwise. A PTM equipped with $p_X$ can perform all basic operations of a Turing machine, as well as determining whether $n \in X$ for any $n$. It can do this by simulating $A_\infty$ in parallel with its main computation, storing the bits of $p$ produced by $A_\infty$ and examining them when needed. If it needs to test whether $n \in X$ and has not yet determined $b_n$, it simply waits until this is found.

In the cases where $p_X$ is a dyadic rational this method will not work, but since $X$ will be recursive, there is a probabilistic Turing machine that can simulate such an $o$-machine without using any probabilistic methods at all. In this way, these methods suffice to simulate any $o$-machine.

Theorem 4.1. For any $o$-machine $M$ with oracle $X$, there is a PTM $P_M$ equipped with probability $p_X$ that when given the same inputs plus one additional input $j$, $P_M$ produces the same output as $M$ with probability greater than $1 - \frac{1}{2^j}$.

Proof. Immediate. □

Since all functions of the form $f : \mathbb{N}^n \rightarrow \mathbb{N}^m$ or $f : \mathbb{N}^n \rightarrow \mathbb{R}^m$ are computable by some $o$-machine, we can see that there are probabilities that would allow PTMs to compute any such functions.

Corollary 4.2. For any function $f : \mathbb{N}^n \rightarrow \mathbb{N}^m$ or $f : \mathbb{N}^n \rightarrow \mathbb{R}^m$, there exists a PTM that when given inputs $j, x_1, \ldots, x_n$ produces $f(x_1, \ldots, x_n)$ with
probability greater than $1 - \frac{1}{2^j}$, produces incorrect output with probability less than $\frac{1}{2^j}$ and diverges with probability 0.

Since these natural and real numbers can be used to code other mathematical objects, this set of PTM computable functions includes a vast number of interesting mathematical functions. Given an appropriately biased coin, a PTM could decide the halting problem or the truths of first order arithmetic.

Finally, just as there is a single universal Turing machine which can take the code of a Turing machine as input and simulate it, so there is a universal $o$-machine which takes the code of an arbitrary $o$-machine and simulates it so long as it is equipped with the oracle of the machine being simulated. A similar job can be performed by a specific PTM, provided that the $o$-machine to be simulated does not have an oracle set that would be encoded as a dyadic rational. As such $o$-machines can only compute recursive functions, this is not a great concern.

**Theorem 4.3.** There is a specific PTM $P_U$ that takes inputs $j, n, m \in \mathbb{N}$ and when equipped with any non-dyadic probability $p_X$, $P_U$ computes the result of applying the $o$-machine with oracle $X$ and index $n$ to the input $m$, producing the correct output with probability at least $1 - \frac{1}{2^j}$.

**Proof.** Immediate. \qed

5 GETTING BY WITH INCREASINGLY ACCURATE BIASES

These same results can all be realised without the need for a coin with an infinitely precise bias. Instead, consider a variant of the PTM which is given a succession of coins $\{c_n\}$ where the $n$-th coin is used for the $n$-th toss. If the probability of $c_n$ coming up heads is given by the rational probability $p_n$ and $\{p_n\}$ converges quickly to some arbitrary real $p$, then all of the above results hold with only minor modifications.

If we once again approximate $p$ using the average number of times heads comes up in $n$ tosses, we find that the mean of $\hat{p}$ is no longer $p$, but $\mu$, where

$$\mu = \frac{\sum_{i=1}^{n} p_i}{n} \leq \frac{\sum_{i=1}^{n} (p + \frac{1}{2^j})}{n} = \frac{np + 1}{n} = p + \frac{1}{n} \quad (5.1)$$

By a similar argument, we find the lower bound for $\mu$, and see that

$$p - \frac{1}{n} < \mu < p + \frac{1}{n} \quad (5.2)$$

The variance is now given by

$$\sigma^2 = \frac{\sum_{i=1}^{n} p_i(1 - p_i)}{n^2} \leq \frac{1}{4n} \quad (5.3)$$
We can now once again use the Chebyshev inequality to form an upper bound for the probability of error. If we set \( n = 2^{j+2k} \) (which is 4 times higher than the value of \( n \) used previously), we see that

\[
P \left( \left| \hat{p} - \mu \right| < \frac{1}{2^{k+1}} \right) \geq 1 - \frac{1}{2^j}
\]

\[
P \left( \left| \hat{p} - p \right| < \frac{1}{2^k} + \frac{1}{2^{j+2k}} \right) \geq 1 - \frac{1}{2^j}
\]

And so this new value of \( n \) suffices in this case. Replacing all later references to \( 2^{j+2k-2} \) with \( 2^{j+2k} \) and references to \( 2^{j+3k-2} \) with \( 2^{j+3k} \), all the theorems follow. It is also easy to see that we could relax our constraint that the sequence of biases converges quickly. Instead it can converge very slowly, so long as there is a recursive bound on how slowly. That way the machine could use this bound to calculate a subsequence of coins whose probabilities would converge quickly.

### 6 GETTING BY WITH RANDOMLY CHOSEN BIASES

Another way that we can avoid the need for a coin with an infinitely accurate bias is via a probability distribution of finitely accurate biases. As in the previous section, we use a sequence of coins \( \{c_n\} \) where the \( n \)-th coin is used for the \( n \)-th toss. This time however, the bias on each coin will be chosen with an independent random trial from a fixed probability distribution. We will see that so long as the mean of this distribution is a non-recursive real, access to this randomisation extends the PTM’s powers. Specifically, it can compute the binary expansion of the mean with arbitrarily high confidence.

We first consider the case of a discrete probability distribution, where the probability of choosing the bias \( x_i \in [0, 1] \) is denoted by \( P(x_i \text{ is chosen}) \). To generate the value of the \( n \)-th coin toss, we must combine the process of randomly choosing a bias with the process of tossing a coin with that bias. Let \( z \) be a random variable representing the result of the coin toss, equaling 1 if the coin lands heads and 0 if tails. From the rules of conditional probability,

\[
P(z = 1) = \sum_i P(z = 1 \mid x_i \text{ is chosen}) P(x_i \text{ is chosen})
\]

\[
= \sum_i x_i P(x_i \text{ is chosen})
\]

\[
= \mu_x
\]

(6.1)
The same is true if we use a continuous distribution \( \rho(x) \). In this case

\[
P(z = 1) = \int_0^1 P(z = 1 \mid x \text{ is chosen}) \rho(x \text{ is chosen}) \, dx
= \int_0^1 x \rho(x \text{ is chosen}) \, dx
= \mu_x
\]

(6.2)

In either case \( P(z = 0) = 1 - \mu_x \). Thus, the combined process of randomly choosing a bias between 0 and 1 from any distribution with mean \( \mu_x \) and then flipping the appropriate coin is equivalent to the process of flipping a single coin with bias \( \mu_x \). From this it is clear that one can determine the binary expansion of \( \mu_x \) with arbitrary confidence using the methods of Sections 2–4. Indeed, all the results of those sections hold for this modified type of PTM without the need for any additional coin tosses.

In the case of discrete distributions, it is interesting to consider how \( \mu_x \) could be non-recursive. Recall that for a discrete distribution, the mean is defined by \( \sum_i x_i p(x_i) \) and that the recursive reals are closed under finite sums and products. Thus, if the distribution is finite, the only possibilities are that at least one of the possible biases is non-recursive or at least one of the associated probabilities is non-recursive. For infinite discrete distributions there is the additional possibility of one or both of the sequences \( \{x_i\} \) and \( \{p(x_i)\} \) being non-recursive despite all the individual elements being recursive.

Therefore, this method of randomly choosing a bias from a given distribution and then flipping a coin with that bias allows a PTM to exceed the power of a Turing machine without relying upon coins with infinitely precise biases.

7 CONCLUSIONS

Over the course of this paper, we have shown three ways to implement an abstract oracle by tossing biased coins. This was achieved by demonstrating a computational equivalence between \( o \)-machines and three different classes of PTM. These results show that it is very careless to say that randomness does not increase the power of the Turing machine. While this is true of fair coins and recursively biased coins, they form only a set of measure zero in the space of all possible biased coins. Indeed, if a bias is chosen completely at random (from a uniform distribution over \([0, 1]\)) then with probability one, it would be non-recursive and thus extend the powers of a Turing machine that had access to it.

This is not only of mathematical interest, but is particularly significant in the study of what is physically computable. There has been continued interest over the years about whether some form of \( o \)-machine might be physically
realisable [2, 4, 7]. A simple way to go about implementing an oracle would be to measure some quantity, such as the distance between two particles, with finer and finer accuracy. If this distance happened to be a non-recursive real, we could then use the methods of section 3 to compute the binary expansion and use this as an oracle set. However, such methods based on measuring continuous quantities could quickly run into fundamental limits of quantum mechanics, especially if there exists some fundamental lengthscale such as that of the Planck scale which is demanded by some theories of quantum gravity.

Using randomness provides an alternative that does not run afoul of these limitations. It allows one to measure an underlying continuous quantity with a sequence of discrete measurements that do not individually become increasingly accurate. It is the increasing total number of measurements that provides the accuracy, so no particular measurement needs to be more accurate than the quantum limits.

In fact, quantum mechanics even suggests a way to simulate such biased coin tosses. A qubit is any quantum system that has two possible states and, when measured, is seen to take on one of these states randomly [6]. Each state of a qubit has an associated probability amplitude, which is a complex number that defines the probability that the system will be found in that state. These probability amplitudes are allowed to be arbitrary complex numbers having moduli less than one and thus the induced probabilities, which are squares of the moduli, are arbitrary reals between 0 and 1. A qubit therefore seems to be a physical implementation of an arbitrarily biased coin.

There is, however, an important difference: while a biased coin can be flipped as many times as one wishes, a qubit is destroyed once its state is determined. Furthermore, by the no cloning theorem of quantum mechanics [11], we cannot get around this destructive measurement by making perfect copies of the qubit.

However, the technique of section 6 seems to offer a way out. If there is any method which creates qubits with biases chosen randomly around some non-recursive mean, then this method implements a non-recursive oracle. Since the non-recursive values this mean could take form a set of measure one in the space of all possible biases, this appears quite plausible and it would seem to require an independent physical principle to force all such methods to pick out only recursive means.

If we furthermore wish to harness this non-recursive power to compute some particular non-recursive function, we need to know more about the non-recursive mean around which our biases are generated. For instance, we could try to create a PTM for deciding whether a given formula of the predicate calculus is a tautology by using a mean that codes the set of halting Turing machines, or even by using the halting probability $\Omega_1$, described by Chaitin [1], in the setting up of a qubit [5].

However, it appears to be very difficult to generate biased qubits around such a known mean. Consider some controllable variable $\lambda$ (such as the amount
of time an electron is exposed to a magnetic field) involved in creating the probability amplitude $a(\lambda)$ for a qubit state and let us suppose that we could generate this controllable variable in some distribution $P_\lambda$, with the appropriate mean. Even then, we would still have further difficulties to overcome as the relationship between the bias of a qubit, represented by $|a(\lambda)|^2$, and the controllable variables that determine it is inevitably non-linear. It is then not sufficient to control the mean of the controlled variables $\lambda$: we must also precisely determine the details both of their distributions $P_\lambda$ and of the functions $a(\lambda)$, which can and will be affected by generally uncontrollable quantum decoherence, to obtain the mean of the quantum probabilities through their distributions $P_a$.

$$P_a = P_\lambda \left/ \left| \frac{d|a|^2}{d\lambda} \right| \right.,$$

(7.1)

and it seems quite unlikely that all of these would be possible.

The use of biased coins to compute more than the Turing machine is certainly of physical interest and, although it is not yet clear how it could be physically harnessed to increase our computational abilities, the close connections with quantum theory suggest a potential for further study.

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