Some misunderstandings in psychology about confidence intervals

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Introduction

Publication bias (e.g., Legget et al., 2013; Francis, 2014) and the replication crisis (e.g., Open Science Collaboration, 2015) in empirical Psychological studies have been discussed especially intensively during the last 10 years (see Nelson et al., 2018 for a review). Many of these empirical studies report results that are too good to be true and the percentage of these studies in Psychology that can be successfully replicated has been estimated to be low.

These problems can be partly attributed to misunderstandings in Statistics. Cassidy et al. (2019) pointed out that the correct definition of a p-value is described only in 11% of Introduction-to-Psychology textbooks (see also Gigerenzer, 2018; Lakens, 2021). The assumptions underlying the central-limit theorem are often missing in the textbooks and the effectiveness of the central-limit theorem for securing normality of a distribution can be overestimated in Psychology (Hesterberg, 2008; Sawada, 2021). The standard error of the mean (SEM), instead of the standard deviation (SD), has been misused to describe variability across samples in Anesthesiology (Nagele, 2003). It seems likely that the same misuse can also be observed in Psychology.

Nowadays, authors of empirical Psychological studies are encouraged to report details of the Statistical tests and their results, including their confidence intervals (CI) when they report results of Psychological experiments (American Psychological Association, 2019). The CI of an unknown parameter is an interval estimate of the parameter. Consider repeatedly conducting a session during which you collect data and compute the CI of this unknown parameter with 100(1 – α)% (e.g., 95% when α = 0.05) of the confidence level. Note that the computed CIs change randomly across the sessions while the unknown parameter is regarded as a constant.

With this done, a CI in each session will cover the parameter with a probability 100(1 – α)% (Figure 1A)1. So, 100(1 – α)% of the CIs estimated in the sessions include the unknown parameter asymptotically when the number of the sessions goes to positive infinity. The computation of the CI of an unknown parameter

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1The CI is often described as a range of values within which the unknown parameter falls with 100(1 – α)% of a probability. This description can lead to a misunderstanding, namely, that the unknown parameter is a random variable and that the CI of the parameter is fixed. This misunderstanding is especially problematical when multiple CIs of a single unknown parameter are discussed (e.g., a forest plot of a meta-analysis).
In this study, we deliberately concentrated our discussion on the CI of the mean of the population because this is the most common CI used in Psychology, as represented in the Statistical textbooks that are currently being used in this field (these textbooks and their editions are listed in our Supplemental Material, https://osf.io/r7b8t/). We surveyed descriptions of the definition and of the equations used to compute the CI in the Statistics textbooks often used in Psychology. We found that the equation in 25% (5/20) of the textbooks computes the CI differently from the way the CI is computed in the other textbooks.

The CI of the mean of a population

Consider a population that is normally distributed. The CI of the mean of this population is computed as:

\[
\left[ \hat{\mu} - T^{-1} (1 - \frac{\alpha}{2}, n - 1) \frac{\hat{\sigma}}{\sqrt{n}}, \hat{\mu} + T^{-1} (1 - \frac{\alpha}{2}, n - 1) \frac{\hat{\sigma}}{\sqrt{n}} \right]
\]

(1)

where \( n \) is the sample size, \( \hat{\mu} \) is the estimated mean of the population, \( \hat{\sigma} \) is the estimated standard deviation and, \( \frac{\hat{\sigma}}{\sqrt{n}} \) is the estimated standard error of the mean, and \( T^{-1}(1 - \frac{\alpha}{2}, n - 1) \) is the inverse cumulative distribution function of the \( t \)-distribution with the degree of freedom \( n - 1 \) and for the probability \( 1 - \alpha/2 \).

The level of the CI is \( 100(1 - \alpha)\% \) (e.g., \( \alpha = 0.05 \) for the 95% CI). Note that the 95% CI of the mean is the range of the value from which the estimated mean \( \hat{\mu} \) is not significantly different with the \( 100\alpha \% \) level \( (p \geq \alpha) \) when a single one-sample, two-tailed \( t \)-test is conducted.

Note that the function \( T^{-1}(1 - \frac{\alpha}{2}, n - 1) \) can be well-approximated with the inverse cumulative distribution function of the standard normal distribution when the degree of freedom \( n - 1 \) is large enough (say \( n \geq 30 \), see also Hesterberg, 2008). The inverse cumulative distribution function of the standard normal distribution is independent from \( n \). If one assumes that the sample size \( n \) is large enough, the \( 100(1 - \alpha)\% \) CI can be written as:

\[
\left[ \hat{\mu} - \Phi^{-1} (1 - \frac{\alpha}{2}) \frac{\hat{\sigma}}{\sqrt{n}}, \hat{\mu} + \Phi^{-1} (1 - \frac{\alpha}{2}) \frac{\hat{\sigma}}{\sqrt{n}} \right]
\]

(2)

2. Consider comparing the means of two populations using their CIs (Equation 1). Goodwin (2010) stated that, if the CI of the mean of one group does not overlap with the CI of the other group, the difference is necessarily significant \((p < \alpha)\) on the basis of an unpaired \( t \)-test (either equal or unequal variances). This statement is true but its inverse is not necessarily true. If the CIs of the means do overlap, the difference may or may not be significant (Schenker and Gentleman, 2001, see Cousineau et al., 2021 for adjusting the CI to compare the means of two populations).
where \( \Phi^{-1}(1 - \frac{\alpha}{2}) \) is the inverse cumulative distribution function of the standard normal distribution for the probability 1 - \( \alpha \)/2. Note that \( \Phi^{-1}(1 - \frac{\alpha}{2}) \) is \( \approx 1.96 \) when \( \alpha \) is 0.05. This number 1.96, alone, is occasionally introduced as a constant to compute the 95% CI in Psychology (e.g., Marks and Yardley, 2004).

The difference between Equations (1) and (2) is attributed to the distribution of the estimator (the estimated mean \( \bar{\mu} \) of the population). Equation (1) is derived by using the \( t \)-distribution, while Equation (2) approximates the \( t \)-distribution with the standard normal distribution. This change of the distributions occurs because the standard deviation of the population is underestimated by the sample standard deviation \( \hat{\sigma} \) computed from the samples when the sample size \( n \) is not sufficiently large. This bias of the sample standard deviation distorts the shape of the distribution of \( \bar{\mu} \) from a normal distribution. If the sample size \( n \) is not sufficiently large, Equation (2) will underestimate the CI (Figure 1B). Both Equations (1) and (2) assume that the population is normally distributed in order to derive the distribution of \( \bar{\mu} \). So, one must know the distribution of the estimator to compute the CIs.

Crucially, using Equation (2) instead of (1) means that the CI computed is asymptotic. In other words, it is only valid when the population is infinitely large. This means that the intuition we presented in the Introduction is no longer valid. When we repeat sessions, we have no guarantee that 100(1 - \( \alpha \))% of the intervals computed with equation (2) contain the true mean. A similar point can be made when the Central Limit Theorem (CLT) is used to circumvent the Gaussian assumption on the population.

Now, consider a population that is not normally distributed and its variance is a non-zero finite value. The CI of the mean of this population depends on the shape of the distribution of the population. According to the Central Limit Theorem, the CI can be computed by using Equation (2) asymptotically when the sample size goes to positive infinity (Dekking et al., 2005; Kwak and Kim, 2017). But, note that in a real experiment, the sample size will always be finite. The CI of the mean can be computed approximately by using Equation (1) when the sample size is finite and sufficiently large. Note that a “sufficiently-large” sample size depends on the shape of the distribution of the population (Cuadras, 2002; Wilcox, 2012). Mathematical bounds can be derived using Berry Essen’s Theorem, and for some reason, the number 30 has become stuck in the literature (Hesterberg, 2008). Note, however, that more regularity through moment conditions is assumed, but hardly ever checked in practice.

If the population is not normally distributed and the sample size is small, the CIs computed using Equations (1) and (2) do not have any theoretically-valid meaning. For a non-normal distribution with a small sample size, the CI of the mean of the population is (Devore, 2011):

\[
\left[ \bar{\mu} - \frac{\bar{\mu} - \mu}{\sqrt{\frac{\sigma^2}{n}}} \right], \left[ \bar{\mu} + \frac{\bar{\mu} - \mu}{\sqrt{\frac{\sigma^2}{n}}} \right]
\]

where \( \mu \) is the true mean of the population and \( \frac{\bar{\mu} - \mu}{\sqrt{\frac{\sigma^2}{n}}} \), is the inverse cumulative distribution function of the distribution of \( \bar{\mu} - \mu \). Note that Equation (3) is theoretically valid but is not practically useful because the inverse cumulative distribution function \( \frac{\bar{\mu} - \mu}{\sqrt{\frac{\sigma^2}{n}}} \) of \( \bar{\mu} - \mu \) is generally unknown. There exists techniques for computing the CI from samples even when \( \frac{\bar{\mu} - \mu}{\sqrt{\frac{\sigma^2}{n}}} \) is unknown (Rousselet et al., 2019) but the techniques rely on Bootstrap. These techniques need a sufficiently large sample size so that the empirical distribution is already close to the actual distribution of the population (Hesterberg, 2015).

We surveyed the equations used to compute the CI in the 20 Statistics textbooks often used in Psychology (these 20 textbooks and their editions are listed in our Supplemental Material, https://osf.io/r7b8t/). Five of the 20 textbooks were personally owned by the first author (TS). The other 15 textbooks were taken from the library at HSE University. Textbooks that do not describe any equation used to compute the CI were not included. We found that 75% (15/20) of the textbooks used Equation (1) that was based on a \( t \)-distribution to compute the CI. Ten of these 15 textbooks also described Equation (2) that is based on the normal distribution for a sample size that was sufficiently large. The remaining 25% (5/20) of the textbooks only used Equation (2).

**Discussion**

In this study, we surveyed descriptions of the computation of the CI of the mean of a population in the 20 Statistics textbooks that are often used in Psychology today. The computation of the CI described in 75% (15/20) of the textbooks was based on a \( t \)-distribution (Equation 1) and the remaining 25% (5/20) of the textbooks used a different computation that was based on a normal distribution (Equation 2). The CIs computed by using these two different methods are different from one another. This inconsistency in the computation of the CI makes it difficult to compare results across studies quantitatively (e.g., meta-analysis).

Nowadays, psychologists are encouraged to report the results of their experiments with CIs (Wilkinson and Task Force on Statistical Inference American Psychological Association Science Directorate, 1999; Cumming and Finch, 2005),
Considering the inconsistency across the textbooks that we noted, we encourage psychologists to describe how their CIs were computed (e.g., distributions, equations, or scripts) when they report CIs. Another remedy for this inconsistency is reporting estimated SDs or/and estimated SEMs with their sample sizes. Equations (1) and (2) show that the CI of the mean can be computed from the estimated SD or the estimated SEM when the sample size \( n \) is available. This means that the amount of information included in the CI that was computed from data is, theoretically, the same as the amount of data in the estimated SD as well as in the estimated SEM (see Francis, 2017). These inclusions can elucidate how the CIs were computed. Psychologists also need caution to read empirical studies reporting the CIs especially if the report does not have these inclusions.

It is worth pointing out that the CI in Psychology often refers to the CI of the mean of the population because only the CI of the mean of the population is described in many of the Statistics textbooks often used in Psychology. This reflects a current practice in Psychology that empirical data are discussed only on the basis of the difference of means across conditions. But, this is bad because it can discourage a quantitative discussion of other parameters (e.g., variance, median, and quantile) with their CIs. Note that the CIs of different parameters are computed by using different equations (see Harding et al., 2014 for a review).

Note that the issues of the CI that are discussed in this study are only a part of problems in the way that Psychologists handle statistics. Statistics education and the textbooks used in Psychology should be improved to address these problems.

**Author contributions**

TS contributed to conception and design of the study. TS and OK conducted the survey. LH contributed the theoretical aspects of the study. TS and LH wrote the manuscript. All authors contributed to manuscript revision, read, and approved the submitted version.

**Conflict of interest**

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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**Supplementary material**

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fpsyg.2022.948423/full#supplementary-material

**References**

American Psychological Association (2019). *Publication Manual of the American Psychological Association*, 7th Edn. Washington, DC: American Psychological Association.

Cassidy, S. A., Dimova, R., Giguère, B., Spence, J. R., and Stanley, D. J. (2019). Failing grade: 89% of introduction-to-psychology textbooks that define or explain statistical significance do so incorrectly. *Adv. Methods Practices Psychol. Sci.*, 233–239. doi: 10.1177/2515245919858072

Cousineau, D., Goulet, M. A., and Harding, B. (2021). Summary plots with adjusted error bars: the superb framework with an implementation in R. *Adv. Methods Practices Psychol. Sci.*, 4(3), 158–175. doi: 10.1177/25152459211035109

Cuadras, C. M. (2002). Geometrical understanding of the Cauchy distribution. *Quaderns d‘Estadística i Investigació Operativa*, 26, 283–287.

Cumming, G., and Finch, S. (2005). Inference by eye: confidence intervals and how to read pictures of data. *Am. Psychol.* 60, 170–180. doi: 10.1037/0003-066X.60.2.170

Dekking, F. M., Kraaikamp, C., Lopuhaä, H. P., and Meester, L. E. (2005). *A Modern Introduction to Probability and Statistics: Understanding Why and How*. London: Springer-Verlag.

Devore, J. L. (2011). *Probability and Statistics for Engineering and the Sciences*, 8th Edn. Boston, MA: Cengage Learning.

Francis, G. (2014). The frequency of excess success for articles in Psychological Science. *Psychon. Bull. Rev.* 21, 1180–1187. doi: 10.3758/s13423-014-0601-x

Francis, G. (2017). Equivalent statistics and data interpretation. *Behav. Res. Methods* 49, 1524–1538. doi: 10.3758/s13428-016-0812-3

Gigerenzer, G. (2018). Statistical rituals: the replication delusion and how we got there. *Adv. Meth. Pract. Psychol. Sci.*, 1, 198–218. doi: 10.1177/25152459187771329

Goodwin, C. J. (2019). *Research in Psychology: Methods and Design*. 6th Edn. Hoboken, NJ: John Wiley and Sons.

Greenland, S., Senn, J. I., Rothman, K. J., Carlin, J. B., Poole, C., Goodman, S. N., et al. (2016). Statistical tests, P values, confidence intervals, and power: a guide to misinterpretations. *Eur. J. Epidemiol.* 31, 337–350. doi: 10.1007/s10654-016-0149-3

Harding, B., Tremblay, C., and Cousineau, D. (2014). Standard errors: A review and evaluation of standard error estimators using Monte Carlo simulations. *Quant. Methods Psychol.* 10, 107–123. doi: 10.20982/qmp.10.2.p107

Hesterberg, T. (2008). It's time to retire the "n ≥ 30" rule. *Proc. Am. Stat. Assoc.* (Statistical Computing Section). Available online at: https://research.google/pubs/pub34906/

Hesterberg, T. C. (2015). What teachers should know about the bootstrap: Resampling in the undergraduate statistics curriculum. *Am. Stat.* 69, 371–386. doi: 10.1080/00031305.2015.1089789

Hoekstra, R., Morey, R. D., Rouder, J. N., and Wagenmakers, E. J. (2014). Robust misinterpretation of confidence intervals. *Psychon. Bull. Rev.* 21, 1157–1164. doi: 10.3758/s13423-013-0572-3
Kwak, S. G., and Kim, J. H. (2017). Central limit theorem: the cornerstone of modern statistics. *Kor. J. Anesth.* 70:144. doi: 10.4097/kjae.2017.70.2.144

Lakens, D. (2021). The practical alternative to the *p* value is the correctly used *p* value. *Perspect. Psychol. Sci.* 16, 639–648. doi: 10.1177/1745691620958012

Leggett, N. C., Thomas, N. A., Loetscher, T., and Nicholls, M. E. R. (2013). The life of *p*: “just significant” results are on the rise. *Quart. J. Exp. Psychol.* 66, 2303–2309. doi: 10.1080/17470218.2013.863371

Marks, D. F., and Yardley, L. (Eds.). (2004). *Research Methods for Clinical and Health Psychology.* London: Sage.

Nagele, P. (2003). Misuse of standard error of the mean (SEM) when reporting variability of a sample. A critical evaluation of four anaesthesia journals. *Br. J. Anaesth.* 90, 514–516. doi: 10.1093/bja/aeg087

Nelson, L. D., Simmons, J., and Simonsohn, U. (2018). Psychology’s renaissance. *Annu. Rev. Psychol.* 69, 511–534. doi: 10.1146/annurev-psych-122216-011836

Open Science Collaboration (2015). Estimating the reproducibility of psychological science. *Science* 349:aac4716. doi: 10.1126/science.aac4716

Rousselet, G., Pernet, C., and Wilcox, R. R. (2019). A practical introduction to the bootstrap: a versatile method to make inferences by using data-driven simulations. *PsyArXiv [Preprint]*. doi: 10.31234/osf.io/8h8f7

Sawada, T. (2021). Conditions of the Central Limit Theorem are rarely satisfied in Empirical Psychological studies. *Front. Psychol.* 12:762418. doi: 10.3389/fpsyg.2021.762418

Schenker, N., and Gentleman, J. F. (2001). On judging the significance of differences by examining the overlap between confidence intervals. *Am. Stat.* 55, 182–186. doi: 10.1198/000313001317097960

Wilcox, R. R. (2012). *Introduction to Robust Estimation and Hypothesis Testing.* 3rd Edn. San Diego, CA: Academic Press.

Wilkinson, L., and Task Force on Statistical Inference American Psychological Association Science Directorate (1999). Statistical methods in psychology journals: guidelines and explanations. *Am. Psychol.* 54, 594–604. doi: 10.1037/0003-066X.54.8.594