Measurement of the Bottom-Strange Meson Mixing Phase in the Full CDF Data Set
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Measurement of the Bottom-Strange Meson Mixing Phase in the Full CDF Data Set

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We report a measurement of the bottom-strange meson mixing phase \( \beta_s \) using the time evolution of \( B_s^0 \to J/\psi (\to \mu^+\mu^-) (\to K^+ K^-) \) decays in which the quark-flavor content of the bottom-strange meson is identified at production. This measurement uses the full data set of proton-antiproton collisions at \( \sqrt{s} = 1.96 \) TeV collected by the Collider Detector experiment at the Fermilab Tevatron, corresponding to \( 9.6 \) fb\(^{-1}\) of integrated luminosity. We report confidence regions in the two-dimensional space of \( \beta_s \) and the \( B_s^0 \) decay-width difference \( \Delta \Gamma_s \), and measure \( \beta_s \in [-\pi/2, -1.51] \cup [-0.06, 0.30] \cup [1.26, \pi/2] \) at the 68% confidence level, in agreement with the standard model expectation. Assuming the standard model value of \( \beta_s \), we also determine \( \Delta \Gamma_s = 0.068 \pm 0.026 \) (stat) \( \pm 0.009 \) (syst) ps\(^{-1}\) and the mean \( B_s^0 \) lifetime, \( \tau_s = 1.528 \pm 0.019 \) (stat) \( \pm 0.009 \) (syst) ps, which are consistent and competitive with determinations by other experiments.

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The noninvariance of the physics laws under the simultaneous transformations of parity and charge conjugation (CP violation) is accommodated in the standard model (SM) through the presence of a single irreducible complex phase in the weak-interaction couplings of quarks. A broad class of generic extensions of the SM is expected to naturally introduce additional sources of CP violation that should be observable, making CP–violation studies promising to search for experimental indications of new particles or interactions. Thus far, CP violation has been established in transitions of strange and bottom hadrons, with effects consistent with the SM interpretation [1–3]. Much less information is available for bottom-strange mesons, \( B_s^0 \). Studies of \( B_s^0 - \overline{B_s^0} \) flavor oscillations are unique in that they probe the quark-mixing (Cabibbo-Kobayashi-Maskawa, CKM) matrix element \( V_{ts} \), which directly enters the mixing amplitude. Large non-SM enhancements of the mixing amplitude are excluded by the precise determination of the oscillation amplitude.
frequency in 2006 [4]. However, non-SM particles or couplings involved in the mixing may also increase the size of the observed CP violation by enhancing the mixing phase \( \beta_s = \text{arg}[-(V_{ts}V_{tb}^*)/(V_{cs}V_{cb})]^2 \) [5] with respect to the value expected from the CKM hierarchy, \( \beta^\text{SM} \approx 0.02 \) [2], henceforth referred to as ‘SM expectation’. A non-SM enhancement of \( \beta_s \) would also decrease the size of the decay-width difference between the light and heavy mass eigenstates of the \( B_s^0 \) meson, \( \Delta \Gamma_s = \Gamma_L - \Gamma_H \). The values of the mixing phase and width difference are loosely constrained, and currently the subject of intense experimental activity. The analysis of the time evolution of \( B_s^0 \to J/\psi\phi \) decays provides the most effective determination of \( \beta_s \) and \( \Delta \Gamma_s \) [6]. Assuming negligible contributions from sub-leading decay amplitudes [7], the underlying \( b \to c\bar{s}s \) quark transition is dominated by a single real amplitude, making \( \beta_s \) the sole CP-violating phase observable, through the interference between the amplitudes of decays occurring with and without oscillations.

The first determinations of \( \beta_s \), by the CDF and D0 experiments, suggested a mild deviation from the SM expectation [8]. The interest in this measurement increased further recently, because of the 3.9\( \sigma \) departure from the SM expectation of the dimuon asymmetry observed by D0 in semileptonic decays of \( B_s^0 \) mesons [9], which is tightly correlated with \( \beta_s \), if generated in the \( B_s^0 \) sector [5]. While updated measurements in \( B_s^0 \to J/\psi\phi \) decays [10–13] showed increased consistency with the SM, more precise experimental information is needed for a conclusive interpretation.

In this Letter we report a measurement of \( \beta_s \); \( \Delta \Gamma_s \); the mean lifetime of heavy and light \( B_s^0 \) mass eigenstates, \( \tau_s = 2/(\Gamma_H + \Gamma_L) \); and the angular momentum composition of the signal sample using the final data set collected by the CDF experiment at the Tevatron proton-antiproton collider, corresponding to an integrated luminosity of 9.6 fb\(^{-1}\). The analysis closely follows a previous measurement in a subset of the present data [10], and introduces an improved determination of the sample composition based on a new study of the \( K^+K^- \) and \( J/\psi K^+K^- \) mass distributions.

The CDF II detector is a magnetic spectrometer surrounded by electromagnetic and hadronic calorimeters and muon detectors that has cylindrical geometry with forward-backward symmetry. Charged particle trajectories (tracks) are reconstructed using single- and double-sided silicon microstrip sensors arranged in seven cylindrical layers [14] and an open cell drift chamber with 96 layers of sense wires [15], all immersed in a 1.4 T axial magnetic field. The resolution on the momentum component transverse to the beam, \( p_T \), is \( \sigma_{p_T}/p_T^2 \approx 0.07\% \) (\( p_T \) in GeV/c), corresponding to a mass resolution of our \( B_s^0 \) signal of about 9 MeV/c\(^2\). Muons with \( p_T > 1.5 \text{ GeV}/c \) are detected in multiwire drift chambers [16]. A time-of-flight detector identifies charged particles with \( p_T < 2 \) GeV/c [17], complemented by the ionization-energy-loss measurement in the drift chamber at higher transverse momenta. The combined identification performance corresponds to a separation between charged kaons and pions of approximately two Gaussian standard deviations, nearly constant in the relevant momentum range. Events enriched in \( J/\psi \to \mu^+\mu^- \) decays are recorded using a low-\( p_T \) dimuon online selection (trigger) that requires two oppositely-charged particles reconstructed in the drift chamber matched to muon chamber track segments, with a dimuon mass between 2.7 and 4.0 GeV/c\(^2\).

In the analysis, two tracks matched to muon pairs are required to be consistent with a \( J/\psi \to \mu^+\mu^- \) decay, with dimuon mass 3.04 < \( m_{\mu\mu} \) < 3.14 GeV/c\(^2\). These are combined with another pair of tracks consistent with a \( \phi \to K^+K^- \) decay, 1.009 < \( m_{\phi K} \) < 1.028 GeV/c\(^2\), in a kinematic fit to a common vertex. A dimuon mass constraint to the known \( J/\psi \) mass [1] improves the \( B_s^0 \) mass resolution. An artificial neural network (NN) classifier [10] combines multiple discriminating variables into a single quantity that statistically separates the signal from the dominant background from combinations of real \( J/\psi \) decays with random track pairs and a minor component of random four-track combinations (both collectively referred to as combinatorics). The NN is trained with simulated events for the signal and data from sidebands in \( B_s^0 \) mass, [5.29, 5.31] \( \cup \) [5.42, 5.45] GeV/c\(^2\), for the background. In decreasing order of discriminating power, the input variables to the NN include kinematic quantities, muon and hadron particle identification information, and vertex fit quality parameters.

Figure 1 shows the \( J/\psi K^+K^- \) mass distribution from the final sample of candidates that pass an NN threshold chosen as to maximize the sensitivity to the measurement of \( \beta_s \) [10]. The distribution shows a signal of approximately 11 000 decays, above a fairly constant background dominated by the prompt combinatorial component, and smaller contributions from mis-reconstructed B decays.

We determine the quantities of interest using a fit to the time evolution of bottom-strange mesons. The differences in time evolution of states initially produced as a \( B_s^0 \) or \( B_s^0 \) meson are included in the fit as well as the differences between decays that result in a CP-odd or CP-even combination of the \( J/\psi\phi \) angular momenta. The proper decay time of a \( B_s^0 \) candidate is a fit observable calculated as \( t = M L_{xy}/p_T \), where \( L_{xy} \) is the distance from the primary vertex to the \( B_s^0 \) decay vertex, projected onto the \( B_s^0 \) momentum in the plane transverse to the beam, \( p_T \); and \( M \) is the known mass of the \( B_s^0 \) meson [1]. The proper decay-time uncertainty, \( \sigma_t \), is calculated from the measurement uncertainties in \( L_{xy} \). Because the \( B_s^0 \) meson has spin zero and \( J/\psi \phi \) have spin one, the \( B_s^0 \to J/\psi\phi \) decay involves three possible angular momentum states of the \( J/\psi\phi \) system. These are combined into three polarization amplitudes, longitudinal polarization (\( A_0 \)), and transverse polarization with
FIG. 1: (Color online) Distribution of $J/\psi K^+ K^-$ mass with fit projection overlaid.

kaon tag (SSKT) deduces the signal production flavor by exploiting charge-flavor correlations of the neighboring kaons produced during its fragmentation. The fraction of candidates tagged by a combination of OST algorithms totals $\varepsilon_{\text{OST}} = (92.8 \pm 0.1)\%$. The probability of wrongly-tagging the meson, $w_{\text{OST}}$, is determined per event and calibrated using 82 000 $B^\pm \rightarrow J/\psi(\rightarrow \mu^+\mu^-)K^\pm$ decays fully reconstructed in the same sample as the signal [20]. Because the $B^\pm$ does not oscillate, the OST tag is compared with the actual flavor, known from the charge of the $K^\pm$ meson. A single scale factor that matches the predicted mistag probability to the one observed in data is then determined to be $1.085 \pm 0.035$. The observed averaged dilution, $D_{\text{OSS}} = 1 - 2w_{\text{OST}}$, equals $(12.3 \pm 0.6)\%$ including the scale factor, resulting in a tagging power of $\varepsilon_{\text{OSS}}D_{\text{OSS}}^2 = (1.39 \pm 0.05)\%$. The SSKT algorithms tag a smaller fraction of candidates, $\varepsilon_{\text{SSKT}} = (52.2 \pm 0.7)\%$, with better precision. In the $B_0^\pm \rightarrow J/\psi\phi$ sample an average dilution of $D_{\text{SSKT}} = (25.9 \pm 5.4)\%$ is achieved including a $0.94 \pm 0.20$ scale factor obtained by measuring the $B_0^\pm$ oscillation amplitude in approximately 11 000 (1 850) $B_0^\pm \rightarrow D_s^- \pi^+(\pi^+\pi^-)$ decays reconstructed in the data corresponding to the first 5.2 fb$^{-1}$ [10]. The resulting SSKT tagging power is $\varepsilon_{\text{SSKT}}D_{\text{SSKT}}^2 = (3.5 \pm 1.4)\%$. Higher instantaneous luminosity conditions in later data resulted in a reduced trigger efficiency for hadronic $B_0^\pm$ decays. Hence, the additional sample of $B_0^\pm \rightarrow D_s^- \pi^+(\pi^+\pi^-)$ decays is too limited for a significant test of the SSKT performance. Because the SSKT calibration is known for early data only, we conservatively restrict its use to the events collected in that period. Simulation shows that this results in a degradation in $\beta_s$ resolution not exceeding 15%.

The unbinned maximum likelihood fit uses 9 observables from each event to determine 32 parameters including $\beta_s$ and $\Delta\Gamma$, other physics parameters such as $B_s^0$ lifetime, amplitudes and phases, and several other quantities, called nuisance parameters, such as tagging dilution scale factors. The fit uses the information of the reconstructed $B_s^0$ candidate mass and its uncertainty, $m$ and $\sigma_m$; the $B_s^0$ candidate proper decay time and its uncertainty, $t$ and $\sigma_t$; the three transversity angles, $\vec{\alpha}$ and tag information, $D$ and $\xi$; where $D$ is the event-specific dilution given by the mistag probability, and $\xi$ is the tag decision. Both tagged and untagged events are used in the fit. The single-event likelihood is described in terms of signal, $P_s$, and background, $P_b$, probability density functions (density henceforth) as

$$L \propto f_s P_s(m|\sigma_m)P_t(t, \vec{\alpha}, \xi|D, \sigma_t)P_{\vec{\alpha}}(\sigma_{\vec{\alpha}})P_b(D) + (1 - f_s)P_b(m)P_b(t|\sigma_t)P_b(\vec{\alpha}|\sigma_{\vec{\alpha}})P_b(D),$$ (1)

where $f_s$ is the fraction of signal events. The signal mass density, $P_s(m|\sigma_m)$, is parametrized as a single Gaussian with a width determined independently for each candidate. The background mass density, $P_b(m)$, is...
parametrized as a straight line. The time and angular dependence of the signal, \( P_s(t, \bar{\rho}, \xi, |D, \sigma_t|) \), for a single flavor tag are written in terms of two densities, \( P \) for \( B_s^0 \) and \( \bar{P} \) for \( B_s^+ \), as
\[
\frac{1 + \xi D}{2} P(t, \bar{\rho}\sigma_t) + \frac{1 - \xi D}{2} \bar{P}(t, \bar{\rho}\sigma_t) \varepsilon(\bar{\rho}),
\]
which is extended to the case of OST and SSKT independent flavor tags. Acceptance effects on the transversity angle distributions are modeled with an empirical three-dimensional joint probability density function extracted from simulation, \( \varepsilon(\bar{\rho}) \), whose largest excursions do not exceed 15%. The time and angular distributions for flavor-tagged \( B_s^0 \) (\( B_s^+ \)) decays, \( P \) (\( \bar{P} \)), are given by the normalized decay rate as functions of decay time and transversity angles of Ref. [19], assuming no CP violation in the decay. Building on previous measurements [21], we model the decay-time density for the background, \( P_b(t|\sigma_t) \), with a \( \delta \)-function at \( t = 0 \), one positive, and two negative exponential functions. All time-dependent terms are convolved with a proper time resolution function, modeled as a sum of two Gaussians with common mean and independent widths determined by the fit. The resulting decay-time resolution is equivalent to that of a Gaussian distribution with 90 fs standard deviation. The background angular probability density is determined from \( B_s^0 \) mass sideband events to factorize as \( P_b(\bar{\rho}) = P_b(\cos \Theta) P_b(\Phi) P_b(\cos \Phi) \). The distributions of the decay-time uncertainty and the event-specific dilution differ for signal and background events, thus their densities are explicitly included in the likelihood. The probability density functions of the decay-time uncertainties, \( P_s(\sigma_t) \) and \( P_b(\sigma_t) \), are described with an empirical model from an independent fit to the data. The signal density, \( P_s(D) \), is determined from binned background-subtracted signal distributions, while the background density, \( P_b(D) \), is modeled from candidates in the signal sidebands. Potential sources of systematic uncertainties, associated with imprecisely known calibration factors of tagging dilutions, are taken into account by floating these factors in the fit within Gaussian constraints.

The likelihood function shows two equivalent global maxima, corresponding to the solutions with positive and negative value of \( \Delta \Gamma_s \), and additional local maxima generated by approximate symmetries [19]. Multiple solutions make the estimation of parameters and their uncertainties challenging with limited sample size. If \( \beta_s \) is fixed to its SM value, the fit shows unbiased estimates and Gaussian uncertainties for \( \Delta \Gamma_s \), \( \tau_s \), polarization amplitudes, and the phase \( \delta_\perp \), yielding
\[
\tau_s = 1.528 \pm 0.019(\text{stat}) \pm 0.009(\text{syst}) \; \text{ps},
\]
\[
\Delta \Gamma_s = 0.068 \pm 0.026(\text{stat}) \pm 0.009(\text{syst}) \; \text{ps}^{-1},
\]
\[
|A_0|^2 = 0.512 \pm 0.012(\text{stat}) \pm 0.018(\text{syst}),
\]
\[
|A_1|^2 = 0.229 \pm 0.010(\text{stat}) \pm 0.014(\text{syst}),
\]
\[
\delta_\perp = 2.79 \pm 0.53(\text{stat}) \pm 0.15(\text{syst}).
\]

The correlation between \( \tau_s \) and \( \Delta \Gamma_s \) is 0.52. We do not report a measurement of \( \delta_\parallel \). The fit determines \( \delta_\parallel \approx \pi \), but the estimate is biased and its uncertainty is non-Gaussian because the likelihood symmetry under the \( \delta_\parallel \rightarrow 2\pi - \delta_\parallel \) transformation [19] results in multiple maxima in the vicinity of \( \delta_\parallel = \pi \). Systematic uncertainties include mismodeling of the signal mass model, decay-time resolution, acceptance description, and angular distribution of the background; an 8% contamination by \( B^0 \rightarrow J/\psi K^*(892)^0 \) and \( B^0 \rightarrow J/\psi K^+\pi^- \) decays misreconstructed as \( B_s^0 \rightarrow J/\psi \phi \) decays; and silicon detector misalignment. For each source, uncertainties are determined by comparing the fit results from simulated samples in which the systematic effect is introduced in the model and samples simulated according to the default model. The uncertainty on the \( \Delta \Gamma_s \) measurement is dominated by the mismodeling of the background decay time. The largest contribution to the uncertainty on \( \tau_s \) is the effect of silicon detector misalignment. The angular acceptance model dominates the systematic uncertainties on the amplitudes.

If \( \beta_s \) is free to float in the fit, tests in statistical trials show that the maximum likelihood estimate is biased for the parameters of interest, and the biases depend on the true values of the parameters. Hence, we determine confidence regions in the \( \beta_s \) and (\( \beta_s, \Delta \Gamma_s \)) spaces by using a profile-likelihood ratio statistic as a \( \chi^2 \) variable and considering all other likelihood variables as nuisance parameters. The profile-likelihood ratio distributions observed in simulations deviate from the expected \( \chi^2 \) distribution, yielding confidence regions that contain the true values of the parameters with lower probability than the nominal confidence level. In addition, the profile-likelihood ratio distribution depends on the true values of the unknown nuisance parameters. We use a large number of statistical trials to derive the profile-likelihood ratio distribution of our data. The effect of nuisance parameters is accounted for by randomly sampling their 30-dimensional space within 5\( \sigma \) of their estimates in data and using the most conservative of the resulting profile-likelihood ratio distributions to derive the final confidence regions. This procedure ensures that the confidence regions have nominal statistical coverage whatever the configuration of nuisance parameters values and increases the size of the \( \beta_s \) confidence interval by about 40%. We determine the confidence level for \( 32 \times 48 \) evenly spaced points in \( \beta_s \in [-\pi/2, \pi/2] \) and \( \Delta \Gamma_s \in [-0.3, 0.3] \; \text{ps}^{-1} \) and smoothly
interpolate between them to obtain a continuous region (Fig. 2). Assuming the standard model values for \( \beta_s \) and \( \Delta \Gamma_s \), the probability to observe a profile-likelihood ratio equal to or higher than observed in data is 54%. By treating \( \Delta \Gamma_s \) as a nuisance parameter, we also obtain \( \beta_s \in [-\pi/2, -1.51] \cup [-0.06, 0.30] \cup [1.26, \pi/2] \) at the 68% C.L., and \( \beta_s \in [-\pi/2, -1.36] \cup [0.21, 0.53] \cup [1.04, \pi/2] \) at the 95% C.L. The fraction of S-wave in the \( K^+K^- \) mass range 1.009–1.028 GeV/\( c^2 \) is determined from the angular information to be consistent with zero with \( \mathcal{O}(2\%) \) uncertainty, which is in agreement with our previous determination [10] and the LHCb and ATLAS results [12, 13], and inconsistent with the D0 determination [11]. An auxiliary simultaneous fit of the \( K^+K^- \) and \( J/\psi K^+K^- \) mass distributions [23] that includes the full resonance structure of the \( B^0 \rightarrow J/\psi K^+\pi^- \) decay [24] is performed. The \( K^+K^- \) mass is fit in a range enlarged to 0.988–1.2 GeV/\( c^2 \) using a relativistic Breit-Wigner for the \( \phi \) meson, the shape suggested in Ref. [25] for the \( f_0(980) \) meson, and an empiric shape determined from data for the combinatorial background. In the 1.009–1.028 GeV/\( c^2 \) mass range, this fit determines a \( (0.8 \pm 0.2(\text{stat}))\% \) \( K^+K^- \) S-wave contribution in agreement with the central fit, and a contamination from mis-identified \( B^0 \) decays of \( (8.0 \pm 0.2(\text{stat}))\% \), which is significantly larger than the 1–2% values typically derived assuming only \( P \)-wave \( B^0 \) decays [10, 11]. If neglected, this additional \( B^0 \) component could mimic a larger \( K^+K^- \) S-wave than present.

In summary we report the final CDF results on the \( B^0 \) mixing phase and decay width difference from the time-evolution of flavor-tagged \( B^0 \rightarrow J/\psi\phi \) decays reconstructed in the full Tevatron Run II data set. This analysis improves and supersedes the previous CDF measurement obtained in a subset of the present data [10]. Considering \( \Delta \Gamma_s \) as a nuisance parameter, and using the recent determination of the sign of \( \Delta \Gamma_s \) [26], we find \(-0.06 < \beta_s < 0.30 \) at the 68% C.L. Assuming a SM value for \( \beta_s \), we also report precise measurements of decay-width difference, \( \Delta \Gamma_s = 0.068 \pm 0.026(\text{stat}) \pm 0.009(\text{syst}) \) ps\(^{-1} \), and mean \( B^0 \) lifetime, \( \tau_s = 1.528 \pm 0.019(\text{stat}) \pm 0.009(\text{syst}) \) ps. All results are consistent with expectations and with determinations of the same quantities from other experiments [11–13], and significantly improve the knowledge of the phenomenology on CP violation in \( B^0 \) mixing.

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