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Genetic and environmental factors affecting some reproductive traits of Holstein cows in Cuba

A Menendez Buxadera¹, L Dempfle²

¹ Centro de Investigación para el Mejoramiento Animal, Carretera Central Km 21 1/2, Cotorro, Havana, Cuba;
² International Trypanotolerance Center, PMB 14, Banjul, Gambia

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Summary — A total of 226 651 fertility records of dairy cows obtained from 1980 to 1988 was studied in order to determine the environmental and genetic factors affecting the reproductive performance of Holstein cows under Cuban conditions. Only 43.9% of the inseminated females were pregnant at first service; however, for heifers this value was 63.1%. The seasonal variability was higher for heifers and for primiparous than for older lactating cows. The best performance was found from February to April, whereas during the hot and humid summer (July to September) poorer results were obtained. Age at calving or number of calvings was another important environmental source of variation: the earlier the calving the poorer is the next reproductive performance. The genetic analyses were made within calving number with the REML procedure. For heifers (226 sires, 45 575 records) the heritability and the genetic coefficient of variation were: 2.26 and 10.94%, 3.24 and 11.24%, and 3.04 and 6.19% for conception rate (CR), numbers of services per conception (SG) and conception status (CS = 1/SG), respectively. For first calving females (280 sires, 43 647 records) the results were: 1.94 and 15.93%, 3.25 and 12.80%, and 3.47 and 9.47% for CR, SG and CS, respectively. For the second and third calving, the results were poorer. For the calving interval and days open, the heritabilities were between 1.86 and 4.64%. The results of SG were selected as the best and more useful traits showing high genetic correlations (> 0.60) for the same traits in different calving number.

Holstein cattle breed / reproduction traits / genetic parameters / tropical conditions

Résumé — Facteurs génétiques et environnementaux intervenant sur les performances de reproduction des vaches Holstein à Cuba. Au total 226 651 enregistrements de fertilité de vaches laitières obtenus de 1980 à 1988 sont analysés en vue de déterminer les

Correspondence and reprints: F Ménissier, Station de génétique quantitative et appliquée, Inra, 78352 Jouy-en-Josas cedex, France
facteurs génétiques et environnementaux intervenant sur les performances de reproduction de vaches Holstein dans les conditions cubaines. Seulement 43,9 % des femelles inséminées sont gestantes à la première insémination, alors que pour les génisses ce taux est de 63,1 %. La variabilité due à la saison apparaît plus élevée chez les génisses et primipares que chez les vaches plus âgées. Les meilleurs résultats sont obtenus de février à avril, alors que les plus mauvais sont observés durant l’été (juillet à septembre), qui est chaud et humide dans les conditions cubaines. L’âge au vêlage et le rang de vêlage sont aussi d’importantes sources environnementales de variation. Plus le vêlage est précoce, plus la reproduction suivante est mauvaise. Les analyses génétiques ont été réalisées intrarang de vêlage en utilisant une méthode REML. Pour les génisses (226 pères, 45 575 performances), les hérédabilités et coefficients de variation génétiques sont de 2,26 et 10,94 %, de 3,24 et 11,24 % et de 3,04 et 6,19 % respectivement pour le taux de conception (CR), le nombre d’inséminations par fécondation (SG) et l’état de gestation (CS = 1/SG). Pour les primipares (280 pères, 43 647 performances), ces résultats sont de 1,94 et 15,93 %, de 3,25 et 12,80 % et de 3,47 et 9,47 % respectivement pour CR, SG et CS. Pour les vaches au deuxième et troisième vêlages, les estimations sont plutôt inférieures. L’intervalle entre vêlages et la durée entre vêlage et fécondation ont des hérédabilités estimées variant entre 1,86 et 4,64 %. Les résultats relatifs au nombre d’inséminations par gestation (SG) conduisent à considérer ce critère comme le meilleur et le plus efficace pour la sélection, manifestant par ailleurs une forte corrélation génétique (> 0,60) entre rang de vêlage pour ce critère.

race bovine Holstein / performances de reproduction / paramètres génétiques / conditions tropicales génétiques

INTRODUCTION

Milk production is the main cattle goal in Cuba. Great amounts of resources have been used, not only in construction of cattle barns and infrastructure, but also thousands of Holstein females have been imported from Canada. A national breeding plan (NBP) was established in 1964, in which artificial insemination (AI) played an important role in crossbreeding between native Zebu cattle and Holstein sires. The general strategy and some results of NBP are offered in Anonymous (1978) and Prada (1984).

The population of purebred Holstein cows is relatively large (more than 100 000 heads) and it is very important to analyze the general performance of all traits of economic importance in order to improve the national breeding scheme. For milk production and its constituents, an analysis was made by De los Reyes (1985) for several environmental factors. The main genetic parameters were reported by Guerra et al (1987), whereas some aspects on genotypic × environmental interaction and its role in the estimation of sire breeding value were presented by Menéndez Buxadera and Guerra (1981) and Menéndez Buxadera et al (1989).

According to the results previously mentioned, it could be concluded, regardless of the stressing environmental factors affecting dairy traits, that a very important genetic variability exists; therefore a breeding scheme in the Holstein population could be carried out. In this sense, it will be very useful to determine the characteristics of reproductive performance of the Holstein breed in Cuba since this is the most important individual factor affecting cattle production in this country.
For the last 15 years, a great number of papers has been published with respect to genetic aspects of reproductive performance of dairy cattle (Janson, 1980; Hansen et al 1983; Jansen, 1986; Weller 1989). In general terms, there is a consensus that a large number of traits related to reproduction shows low heritability (lower than 10%), but a high genetic variability. Considering this fact and also the economic importance of fertility, certain traits of reproductive performance should be considered in the selection criterion (secondary traits) for a breeding scheme.

Unfortunately, under tropical conditions almost no research related to this topic has been conducted. For this reason, a study of the genetic and environmental aspects of reproductive performance of Holstein cattle in Cuba was carried out and the main results will be presented in this paper.

MATERIALS AND METHODS

Available data

A total of 270,000 individual records of Holstein cows calving from January 1980 to December 1988 throughout the country was available for this study. These data sets belong to a system developed at the National Center for Cattle Recording (CENCOP) to maintain the control of the Holstein as a pure breed but also some important reproductive data have been collected:

- date of pregnancy and calving;
- registration number of the cows and sire of the calf;
- number of insemination services in which the cow was pregnant;
- herd code;
- results of calving:
  - abortion – if gestation period was less than 260 days;
  - stillbirth or not – if gestation period was between 260 and 295 days;
  - sex of the calf – no abortion.

Taking into account the data available, a procedure was conducted in which the permanent file of identification of each animal maintained at CENCOP was merged with the reproductive file mentioned previously in order to obtain the registration number of the sire of the cow and birth date of the cow.

As a consequence of this process, the new data set has a total of 232,291 records with the information of each cow.

The number of calvings was not available in the original file, so it was decided to generate according to age at calving. Table I shows the general characteristics of the definitive data set.

Only Holstein sires born in Cuba and with a minimum of 20 daughters in a specific calving and distributed in at least five herds were considered for genetic studies; however the exact figures will be presented in the respective tables.
The following reproductive traits were studied: 

- number of services per pregnancy \( SG_i \);
- conception status \( CS_i = 100 \cdot (1/SG_i) \);
- rate of conception for first services \( CR_i = (100 \text{ to pregnant at first service, 0 otherwise}) \);
- calving interval (in days), \( CI_i(i: \text{number of calving}) \);
- interval between calving and pregnancy (in days) \( DO_i(i: \text{number of calving}) \).

Records ending with an abortion, as well as records outside of the range 300–730 days for CI and 20–450 days for DO were deleted. The first three \( CI_i \) and \( DO_i \) were analyzed. With the logical exception of \( DO_i \) and \( CI_i \), all reproductive traits are determined for any value of \( i \) (between 0 and 4).

**Management systems**

The organization of cattle production in Cuba is mainly based on large state enterprises (more than 10,000 heads), which are fairly uniform with respect to structure and organization. The dairy units have around 200 females which are milked twice daily with milking machines. From 10 am to 4 pm, all cows are kept indoors where they receive forage or silage and water, and graze the rest of the time. Inseminations are conducted early in the morning or late at night and the service bulls are selected yearly according to a plan for each enterprise. As a rule, between two and five service bulls are used at the same time in each herd. Pregnancy diagnosis is made by rectal palpation (after 60–90 days of insemination, by a specific technician). All the individual records of each animal are maintained at the unit, and once a month an official inspector from CENCOP visits the herd in order to estimate milk yield of each cow in each milking. All data are sent to CENCOP once a month.

**Statistical procedures**

Two different statistical analyses were performed. The first model was as follows:

\[
Y_{ijkl} = \mu + H_i + M_j + A_k + e_{ijkl}
\] [*]
where $\mu$ is the general mean; $H_i$ is a fixed effect of a combination of herd and year which was absorbed; $M_j$ is a fixed effect of month of calving (or month of pregnancy) for $j = 1, \ldots, 12$; $A_k$ is a fixed effect of number of calvings (or age at pregnancy) for $k = 1, \ldots, 4$; $e_{ijkl}$ is a random residual.

This model was applied in order to determine the magnitude of the fixed effects of certain environmental factors affecting each dependent variable. Age at calving was considered to be a factor. Solutions for age and months were estimated after constraining the last level of each factor to zero.

The estimation of genetic parameters was the objective of the second mixed model which has the following matrix notation:

$$Y = X\beta + Zs + e$$  \hspace{1cm} [2]

where: $Y$ is a vector of observations; $\beta$ is a vector of fixed effects including age and herd–year–season of calving (natural trimester); $s$ is a vector of random sire effects and $e$ is a vector of random residuals, and $X$ and $Z$ are incidence matrices.

The following assumptions were made:

$$E(s) = 0 \quad \text{Var}(s) = I\sigma_s^2$$

$$E(e) = 0 \quad \text{Var}(e) = I\sigma_e^2$$

$$E(y) = X\beta \quad \text{Var}(y) = ZZ' \sigma_s^2 + I\sigma_e^2$$

The $\sigma_s^2$ and $\sigma_e^2$ were estimated by restricted maximum likelihood (REML) (Patterson and Thompson, 1971). The general statistical properties and description of this method of variance components are well illustrated by Kennedy (1981), Dempfle et al (1983) and Lin and McAllister (1984). The variance components estimated for SG$_i$ were used for breeding value estimation (EBV) through the BLUP procedure with a mixed model similar to model 2, but with the age at calving as a covariable. The computer program was developed by Caleyo (1989) and was applied within calving number. The relationship matrix was not considered.

The genetic correlation (Rg) between SG$_i$ was estimated by two methods. The first procedure was published by Calo et al (1973) and was very well presented by Blanchard et al (1983) and is based on the weighted covariance between estimated EBV$_i$; in that case, only those sires with more than 70 effective daughters on each calving were considered. The second estimation of Rg was based in a REML procedure with a model similar to model 2 but using the expected components of the variance of a new trait SG$_T$ = (SG$_i$ + SG$_{i+1}$). In order to fulfill these conditions, a new data set was formed with those cows with records on SG$_i$ in adjacent calvings.

RESULTS

The distribution of data for number of services per conception for different calvings is shown in table II. Only 43.86% of the females were pregnant at first service (this result is equivalent to conception rate to first insemination) and gives clear evidence of the low reproductive performance of Holstein cattle under Cuban conditions. The results for heifers are quite acceptable; however, for lactating females, a dramatic breakdown is observed in fertility rate.
Table II. Frequency distribution for number of services per conception in the first four calvings of Cuban Holstein cows.

| Number of inseminations per conception | Number of data | Overall frequency (%) | Frequency (%) within calving |
|---------------------------------------|----------------|-----------------------|----------------------------|
|                                       |                |                      | 0  | 1  | 2  | 3  | ≥ 4 |
| 1                                     | 99 418        | 43.86                 | 63.05 | 41.37 | 37.44 | 37.03 | 35.77 |
| 2                                     | 54 443        | 23.89                 | 21.37 | 24.26 | 24.46 | 25.34 | 25.42 |
| 3                                     | 29 911        | 13.20                 | 8.61  | 13.60 | 14.90 | 15.07 | 15.67 |
| 4                                     | 17 357        | 7.66                  | 3.94  | 8.14  | 9.11  | 9.09  | 9.12  |
| 5                                     | 10 541        | 4.65                  | 1.75  | 5.28  | 5.67  | 5.31  | 5.59  |
| 6                                     | 6 332         | 2.79                  | 0.74  | 3.19  | 3.45  | 3.24  | 3.50  |
| 7                                     | 3 879         | 1.71                  | 0.34  | 2.00  | 2.12  | 1.90  | 2.20  |
| 8                                     | 2 338         | 0.98                  | 0.11  | 1.04  | 1.33  | 1.27  | 1.31  |
| 9                                     | 2 832         | 1.25                  | 0.09  | 1.12  | 1.52  | 1.75  | 2.02  |
| Total                                 | 226 651       | 100                   | 52 879 | 61 331 | 44 023 | 29 262 | 39 156 |

**Environmental effects**

The results of the statistical analysis according to model 1 show very highly significant effects ($P < 0.001$) for the different factors included; however, the determination coefficient ($R^2$) of the model was between 9.2 and 12.4% for all traits evaluated in heifers and first calving stages. For older animals (more than two calvings) the $R^2$ was between 3.6 and 26.9%. In all cases, the highest $R^2$ was obtained for SG$_i$ and the lowest for CR$_i$.

The solutions for month effects on heifer, first, second and third calving conception status (CS$_i$) are shown in figure 1. Although only CS$_i$ is presented, the general pattern was the same for the rest of the characters. The best results were obtained in March and April (lower SG$_i$, and higher CR$_i$ and CS$_i$), whereas the poorest fertility rate was in September and October. Plotting climatological data on the same figure shows the very evident relationship between the hot and humid summer and a lower reproductive performance, and the relationship between the winter and the less humid period from January to April and the best results for CS$_i$, CR$_i$ and SG$_i$.

Concerning days open (DO$_i$) and calving interval (CI$_i$), month effects were the same as for CS$_i$, CR$_i$ and SG$_i$, which was as expected. However, the most important component of DO$_i$ and CI$_i$ is the interval between calvings and first service and these characters were not available in our data sets, so no more details can be provided on these traits.

The year effect was highly significant ($P < 0.001$) for the main traits studied. The solutions for each character expressed as deviations from the last year are shown in table III. In general terms, during the period of time represented in our data sets, a positive trend in reproductive performance of our Holstein females was found; however, even with these changes an optimum or near optimum level of fertility is never reached. The analysis of the effect of years was made also within calving number and the same pattern was found except in heifers, which showed an opposite trend during the period.
Fig 1. Solutions for months for conception status (%) for different categories of female Holstein cows in Cuba.

Table III. Solutions for year effects on different reproductive traits\(^a\).

| Year | Number of observations\(^b\) | Traits | CS  | CR  | SG   | DO  | CI  |
|------|------------------------------|--------|-----|-----|------|-----|-----|
| 1980 | 12 178                       | −7.0   | −8.7| 0.45| 62   | 63  |
| 1981 | 14 807                       | −5.2   | −5.2| 0.42| 51   | 51  |
| 1982 | 17 221                       | −4.9   | −6.9| 0.25| 40   | 40  |
| 1983 | 22 771                       | −3.2   | −3.2| 0.1 | 62   | 63  |
| 1984 | 25 960                       | −4.2   | −5.4| 0.24| 33   | 33  |
| 1985 | 35 299                       | 1.1    | 1.0 | −0.09| 32   | 33  |
| 1986 | 30 127                       | −2.4   | −3.7| 0.12| 41   | 41  |
| 1987 | 30 196                       | −4.6   | −6.9| 0.15| 0    | 0   |
| 1988 | 33 212                       | 0      | 0   | 0    | 0    | 0   |
|      | General mean                 | 64.2   | 44.1| 2.32| 150  | 434 |

\(^{a}\) The symbols for each trait and units of expression are: SG, number of services for pregnancy; DO, interval from calving to next pregnancy, (days); CS, conception status (%); CI, calving interval, (days); CR, conception rate for first service, (%). \(^{b}\) The numbers for CI and DO were lower (29%) in these two traits. For obvious reasons, there are no data for 1988.
The effect of age or number of calving on the dependent variables was highly significant \((P < 0.001)\). This was expected according to the information offered in table II. In general, the reproductive performance decreases as the number of calving increases as a consequence of cumulative stress due to lactation and previous reproductive disorders. When data were analyzed within calving number, the trend of age effects was negative, so the younger the calving is reached the poorer is the next reproductive performance of lactating females.

**Genetic effects**

The results of the mixed model show a highly significant effect \((P < 0.001)\) of sires for all traits. The heritability \((h^2)\) for each character and variance component estimated by REML are presented in table IV, where the additive genetic coefficient of variation \((CV_g)\) is included.

For \(CR_i\) and \(CS_i\), the genetic variance was more or less the same for different calvings, with the exception of second calving in which an unexpected result was found. For \(DO_i\) and \(CL_i\), an opposite pattern was found and a clear reduction in the genetic variance was obtained for older cows compared to primiparous cows. The genetic variance for \(SG_i\) increased 2.8 times for first calving with respect to results in heifers; however, for second and third calving the estimates were lower. The total phenotypic variance increases when estimating in heifers to third calving. As a consequence of these particular trends in both variance components, the \(h^2\) for all traits was higher for heifers and first calving. The same results were obtained for \(CV_g\).

Concerning the number of reproductive characters available it would be necessary to perform a certain type of discrimination. In this context, it will be very important to take into account not only the value of \(h^2\) and \(CV_g\), but other peculiarities, such as facilities for recording and multiple objectives in the principal factors limiting the level of productivity of populations. In this sense, it will be very useful to conduct a relative comparison among some fertility parameters in different populations of Holstein cows \(\text{(table V)}\). Before 100 days after calving there is not a clear difference among the populations of percentage of females inseminated, which is very important to consider because the climatological conditions, systems of management and level of feeding are quite different in the three countries. The real problems emerge when we look at the percentage of pregnant females before 100 days after calving. According to this complete relative comparison and taking into account the results of Caral et al \(\text{(1984)}\), who reported that only 43\% of cows not pregnant at first service presented second heat in a normal period of time, it can be concluded that at least in female Holstein populations, limiting factors are closely related to the number of services per conception \((SG_i)\). Together with this advantage, this character is related to first service conception rate, is easy to record under field conditions in heifers and lactating cows and, furthermore, it will be useful for other purposes such as sire fertility evaluation; thus, \(SG_i\) will be the preferred trait in our breeding objectives. However, it is necessary to determine the relationship between \(SG_i\) in different types of females.
The general pattern of this study (table VI) shows that there are medium to high genetic associations between SGi at different calvings. The results were poorer with non-consecutive records in comparison to adjacent ones. The Rg estimated by the covariance of EBVI was higher than the results of REML; however, the differences would not suggest that the same genetic bases exist for this trait measured at different stages of the reproductive life of the cows.

Table IV. Variance components, heritability ($h^2$, %) and genetic coefficient of variation (CVg, %) for some reproductive traits in Cuban Holstein cows.

| Characters$^a$ | Calving number$^b$ | Variance | $h^2$ | CVg % | Number of observations$^c$ |
|---------------|------------------|----------|-------|-------|--------------------------|
|               | Genetic          | Total    |       |       |                          |
| CR            | 0                | 47.8504  | 2.26  | 10.94 | 286                      |
|               | 1                | 42.6093  | 1.94  | 15.93 | 280                      |
|               | 2                | 10.9601  | 0.49  | 11.09 | 226                      |
|               | 3                | 49.9943  | 2.21  | 17.58 | 165                      |
| SG            | 0                | 0.03424  | 3.24  | 11.24 | 45 575                   |
|               | 1                | 0.09500  | 3.25  | 12.80 | 43 647                   |
|               | 2                | 0.04913  | 1.59  | 10.21 | 31 497                   |
|               | 3                | 0.07023  | 2.74  | 11.16 | 20 615                   |
| CS            | 0                | 23.2960  | 3.04  | 6.19  |                          |
|               | 1                | 34.6743  | 3.47  | 9.47  |                          |
|               | 2                | 22.1467  | 2.15  | 8.06  |                          |
|               | 3                | 27.3545  | 2.74  | 9.02  |                          |
| DO            | 1                | 324.1464 | 3.67  | 12.09 | 257                      |
|               | 2                | 170.5853 | 2.41  | 10.19 | 236                      |
|               | 3                | 257.8096 | 2.29  | 10.37 | 177                      |
| CI            | 1                | 326.7983 | 4.64  | 4.26  | 24 814                   |
|               | 2                | 207.6600 | 2.31  | 3.34  | 21 285                   |
|               | 3                | 187.1134 | 1.86  | 3.18  | 14 725                   |

$^a$ The symbols for each trait and units of expression are: SG, number of services for pregnancy; DO, interval from calving to next pregnancy, (days); CS, conception status, (%); CI, calving interval (days); CR, conception rate for first service, (%); CVg, (genetic standard deviation/mean of the trait) x 100. $^b$ It is referred to heifers, first, second and third calvings, respectively. $^c$ The number of sires in each line of CR; on the line of SG, number of data. The same meaning for the numbers presented for DO and CI.

Table V. Percentage of Holstein cows inseminated and pregnant before 100 day after calving in three populations.

| Country     | Number of data | Percentage of cows: | Author                  |
|-------------|----------------|----------------------|-------------------------|
|             |                | Inseminated | Pregnant |              |
| Cuba        | 29 476         | 76         | 37       | Menendez Buxadera (1987*) |
| Holland     | 79 987         | 78         | 73       | De Kruijf (1975)        |
| United States | 19 633      | 80         | 78       | Foote (1978)            |

* Unpublished data.

The general pattern of this study (table VI) shows that there are medium to high genetic associations between SGi at different calvings. The results were poorer with non-consecutive records in comparison to adjacent ones. The Rg estimated by the covariance of EBVI was higher than the results of REML; however, the differences would not suggest that the same genetic bases exist for this trait measured at different stages of the reproductive life of the cows.
DISCUSSION

Unfortunately, there are not many references available on reproductive performance of Holstein cattle in the tropics. However, it is obvious from our results that the level is very low (see table II). According to Roman Ponce (1992) this is the general trend in tropical conditions and this low fertility rate of cattle represents the principal limiting factor affecting the productivity of cattle in these regions.

The seasonal variation obtained in these results is similar to those patterns reported in the subtropical regions of Mexico and Florida (Ingraham et al, 1974; Thatcher et al, 1984) and in the very hot and arid environmental conditions of Arizona (Monty, 1984) and Israel (Heimann, 1982; Ron et al, 1984). Many studies have been conducted in order to study the role of heat stress in the summer on the productive and reproductive performances of cattle in that area, with the objective of reducing the depressive effects with a new management system. According to Thatcher (1974), Thatcher et al (1984), Monty (1984) and Wazdauskas et al (1986), the maximum temperature after insemination is one of the main factors affecting the reproductive performance of Bos taurus cattle in tropical countries. High temperature causes a complete hormonal imbalance, which in turn may change the flow of nutrients to the uterus, and at the same time raises the uterus temperature. This mechanism contributes to creating a more hostile ‘uterus environment’ and increases the probability of fertility failure or subsequent embryo death. In figure 1 it can be observed that the magnitude of the month effect is very intense for all cattle categories; however, it would be evident that this environmental stress was twice as high for heifers and first calving females as for the rest of the cows. In fact, this behavior is contradictory in the literature (Thatcher, 1974; Thatcher et al, 1984), according to which the older lactating females must be more variable and more sensitive to heat stress. It is possible to speculate on this apparent paradox. When an adequate level of feeding and management is available, the live weight at first calving of a medium Holstein female is 500 kg, ie, 80% of the adult weight. Under
such conditions, the variability of fertility traits in heifers and first calving females can be less affected by heat stress; however, this is very far from our conditions. The development and body weight curves of Holstein females in Cuba are poorer than those expected for this breed (Menéndez Buxadera et al, 1983); on average our females reach approximately 400 kg live weight in the middle of first calving, which means that this type of animal is less mature (in the sense of weight). Therefore, a higher percentage of the nutrients received must be used for the growing process. This mechanism could account for the higher variability and sensitivity of younger females under our conditions.

All these environmental factors require a more detailed and profound research in order to determine the physiological mechanism. However, the practical consequences of the seasonal variation in reproductive performance of Holstein females under our tropical conditions could be used. For instance, these results show that from September to January around 60% of the cows (not heifers) calving in that period corresponded to the season of higher fertility rate (see fig 1). According to this pattern, it would be highly recommended to inseminate the majority of the females from January to April with semen of Holstein sires of higher breeding value for milk production (principal traits in our breeding scheme). The preliminary estimates of such a policy show an improvement of not less than 20% in genetic potential in our female population with respect to the normal practice of bulls in service all year round.

According to our results it was clear that the $h^2$ for different fertility traits were low (< 5%) and similar to many reports on this subject (see review of Menéndez Buxadera, 1993); however, the genetic variability (table IV) was higher than most of the results in the literature (see Janson, 1980; Berger et al, 1981; Raheja et al, 1989 for SG and Baptist and Gravert, 1973; Janson, 1980; Jansen, 1986 for non-return rate as a trait related to CR). This is due to the reduced culling rate applied in our female population, so these data show a complete picture of the real genetic variability in the reproductive performance of Holstein cattle under the tropical conditions of Cuba. The potential of this variability has not been considered by many researchers on account of the low $h^2$ value. According to Maijala (1987) this is a pessimistic point of view regarding the possibilities of genetic improvement for fertility traits.

For DO$_i$ and CI$_i$ there was a clear reducing $h^2$ pattern when the number of calving increased. In fact, this trend is contrary to the results published by Berger et al (1981), Hansen et al (1983), Jansen (1986) and Strandberg and Danell (1988) who obtained a small increase in $h^2$ and CVg in older females. For the rest of the traits it was evident that these genetic parameters were higher in the first two stages (heiifer and first calving), which is in correspondence to Van Raden et al (1987), Raheja et al (1989) and Weller (1989). The results for the second calving were quite low and no clear explanation was found.

As a matter of fact, the general pattern for higher variability in younger females in comparison to multiparous cows could be a very important advantage in our breeding scheme, since the EBV for characters of the reproductive complex must be determined as early as possible, and higher genetic variability is equivalent to major possibilities for selection and breeding. This is true only if the traits evaluated in heifers and first calving females were in close relationship with the same characters in older cows.
The information available on that subject is contradictory. European reports (Janson, 1980; Distl, 1982; Jansen, 1986) showed that the same reproductive characters in both types of females present the same genetic base so the results of heifers and primiparous cows can be used for sire evaluations. Another recommendation was made by North American researchers (Hansen et al., 1983; Raheja et al., 1989; Van Raden et al., 1989) who reported lower genetic correlations \( (R_g < 0.30) \) although the tendency was the same.

According to our results (table VI) the genetic correlation \( (R_g) \) for \( SG_i \) in different calvings was higher than 0.60 when \( R_g \) was estimated by the weighted covariance between EBV of the sire, whereas \( R_g > 0.49 \) when the REML procedure was used. Both estimates are difficult to compare because the value of \( R_g \) depends on the quality of the data set. However, great differences were not observed. Our point of view is in favor of the results obtained through the EBV of the sire, because in REML the same design matrix was used and the computing requirement was greater.

According to our results it is possible to conclude that the same genetic bases exist between certain fertility traits measured in different stages of the reproductive life of cows. This aspect will be very important for sire evaluation purposes since the records of heifers and first calving could be used in this sense.

**CONCLUSIONS**

This study shows positive yearly trends in the level of reproductive performance of Holstein cattle under Cuban conditions, nevertheless, the general mean was quite low. Calving number and month of calving were the principal environmental factors affecting the fertility rate of Holstein females; this aspect was more critical in heifers and primiparous in comparison to older lactating cows. According to these results the best performance was achieved from February to April, corresponding to the winter season, while during the hot and humid summer (July to September) very poor results were obtained. This seasonal pattern must be interpreted as a general reproductive behavior of Holstein cattle in the tropics, which could be used under practical conditions.

From a breeding point of view, our study confirms that the \( h^2 \) of the characters corresponding to the reproductive complex shows the expected low value \( (h^2 < 5\%) \). However, the genetic coefficient of variation was higher than the value reported for developed and temperate countries (see Menendez Buxadera, 1993), so plenty of space is available for breeding. The number of services per conception \( (SG_i) \) was selected as the character more related to our limiting factor affecting the reproductive performance of this breed under tropical conditions. Furthermore, \( SG_i \) could be used for various purposes and it will be possible to measure this at an early age (heifers). In this sense, the relationship between \( SG_i \) in heifers and first calving females was 0.60 and higher than the 0.88 found between older cows, which allows us to use all information available for breeding value estimation of AI sires. Taking into account that the procedures used for \( R_g \) estimation could be affected by some bias, it is highly recommended to use a more powerful statistical procedure such as an individual multitrait animal model, with different matrix design, although the computing demand will be higher.
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