Effect of angiotensin-converting enzyme inhibitors and angiotensin II receptor blockers on cardiovascular events in patients with heart failure: a meta-analysis of randomized controlled trials

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Abstract

Background: Heart failure (HF) remains a significant cause of morbidity and mortality. Multiple trials over the past several years have examined the effects of both angiotensin-converting enzyme inhibitors (ACEIs) and angiotensin II receptor blockers (ARBs) in the treatment of left ventricular dysfunction, both acutely after myocardial infarction and in chronic heart failure. Yet, there is still confusion regarding the relative efficacy of rennin-angiotensin-aldosterone system (RAAS) inhibition. Our study was conducted to assess efficacy of ACEIs and ARBs in reducing all-cause and cardiovascular mortality in heart failure patients.

Methods: We included randomized clinical trials compared ACEIs and ARBs treatment (any dose or type) with placebo treatment, no treatment, or other anti-HF drugs treatment, reporting cardiovascular or total mortality with an observation period of at least 12 months. Data sources included Pubmed, EMBASE, the Cochrane Central Register of Controlled Trials. Dichotomous outcome data from individual trials were analyzed using the risk ratio measure and its 95%CI with random-effects/ fixed-effects models. We performed meta-regression analyses to identify sources of heterogeneity. All-cause mortality and CV mortality were thought to be the main outcomes.

Results: A total of 47,662 subjects were included with a mean/median follow-up ranged from 12 weeks to 4.5 years. Of all 38 studies, 32 compared ACEIs with control therapy (included 13 arms that compared ACEIs with placebo, 10 arms in which the comparator was active treatment and 9 arms that compared ACEIs with ARBs), and six studies compared ARBs with placebo. ACEIs treatment in patients with HF reduced all-cause mortality to 11% (risk ratio (RR): 0.89, 95% confidence interval (CI): 0.83–0.96, p = 0.001) and the corresponding value for cardiovascular mortality was 14% (RR: 0.86, 95% CI: 0.78–0.94, p = 0.001). However, ARBs had no beneficial effect on reducing all-cause and cardiovascular mortality. In head-to-head analysis, ACEIs was not superior to ARBs for all-cause mortality and cardiovascular deaths.

Conclusions: In HF patients, ACEIs, but not ARBs reduced all-cause mortality and cardiovascular deaths. Thus, ACEIs should be considered as first-line therapy to limit excess mortality and morbidity in this population.

Keywords: Heart failure, ACEIs, ARBs, Meta-analysis, Mortality

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Background
Chronic heart failure (HF) has one of the highest morbidity and mortality rates for cardiovascular diseases worldwide, which affects 1–2% of the adult population in developed countries [1]. To lower the risk of adverse clinical outcomes is therefore extremely important in the therapy of this chronic disease.

It is generally accepted that one of the pathophysiological mechanisms of heart failure is excess activation of the rennin angiotensin aldosterone system (RAAS), so that blockade of the RAAS is one of the key therapeutic targets in patients with HF [2–6]. Recent years, a lot of clinical trials have confirmed that suppression of RAAS (angiotensin-converting enzyme inhibitors (ACEIs) and angiotensin II receptor blockers (ARBs)) reduces cardiovascular (CV) events in patients with heart failure [7–13].

Moreover, the cardioprotective effects of RAAS were recently called into question. The SOLVD study [5] demonstrated that the addition of enalapril to conventional therapy significantly reduced mortality and hospitalization due to heart failure in HF patients. In the ELITE study [14], it was found that treatment with losartan was associated with lower all-cause mortality than captopril. But, in several head-to-head trials (such as the ELITE II study, the VALIANT study, the RESOLVD study and the OPTIMAAL study), ARBs did not significantly reduce cardiovascular mortality as compared with ACEIs [9–12].

Recent meta-analysis reported that in HF patients with hypertension [15] and diabetes [16], treatment with ACEIs resulted in a significant further reduction in all-cause and CV mortality, whereas ARBs had no benefit on these outcomes. These studies indicate that there are different clinical outcomes between ACEIs and ARBs among patients with heart failure.

In light of these conflicting reports, the present meta-analysis was conducted to assess the efficacy of ACEIs and ARBs on all-cause and CV mortality in patients with heart failure.

Methods
Literature search
We searched the database through PubMed, EMBASE, and the Cochrane Central Register of Controlled Trials for randomized clinical trials (RCTs) from November 1977 to June 2017 using Medical Subject Heading ‘anti-hypertensive agents’ or ‘angiotensin II type 1 receptor blockers’ or ‘angiotensin-converting enzyme inhibitors’ and ‘heart failure’. Additionally, studies in the reference lists of the identified articles were also hand searched. The search was limited to RCTs, human subjects and English. The process was strict to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) statement [17].

Study eligibility
Studies were deemed eligible if they: 1) were RCTs, targeting HF patients with reduced ejection fraction (HFrEF, left ventricular ejection fraction ≤45%), with a median or mean follow-up of more than 12 months; 2) compared ACEIs and ARBs treatment (any dose or type) with placebo treatment, no treatment, or other anti-HF drugs treatment; 3) reported cardiovascular or total mortality. When the outcomes obtained from the same population in different publications, only the latest report was included in the analysis.

Data extraction and quality assessment
Two independent investigators (Y. X. and D. X.) extracted data from these reports, and disagreements were resolved by consensus. After excluding the unrelated studies, the following data were extracted: study characteristics (author, publication year, sample size, follow-up period), population baseline characteristics (age, sex, cause of heart failure, risk factors) and end-points. Study quality was assessed using the Jadad score, which is a five-point quality scale, with low quality studies having a score of <2 and high-quality studies a score of ≥3 [18].

Endpoint
All-cause mortality and CV mortality were thought to be the main outcomes.

Statistical analysis
Dichotomous outcome data from individual trials was analyzed using the risk ratio (RR) measure and its 95% confidence interval (CI) [19]. Overall effect was estimated using the Mantel-Haenszel method for RRs [20, 21].

Heterogeneity was evaluated using χ² tests and I² statistics. Studies were considered statistically heterogeneous if I² > 50% and p ≤ 0.05. If heterogeneity between studies were identified, a random-effects model was applied. Otherwise, a fixed-effects model was taken instead [22]. Publication bias was assessed with funnel plots and the Begg regression test [22].

In sensitivity analysis, we removed anyone of the study at a time and repeated the meta-analysis to ensure that no single study would be responsible for the significance of any result separately [22].

Meta-regression was conducted to explore the potential heterogeneity related to the participants (age, cause of HF, left ventricular ejection fraction, and follow-up weeks), the agent used (different types). P < 0.05 was considered statistically significant [22].
Meta-analysis was performed by the Review Manager software (Version 5.0. Copenhagen: The Nordic Cochrane Centre, the Cochrane Collaboration) and the Stata software (version 12.0; Stata Corporation, College Station, TX).

Results
Eligible studies and baseline characteristics
Initial search identified 1002 reference articles, of these 107 relevant articles were selected and reviewed. Then, several studies were further excluded because they were publications from the same trial (n = 7) or reported of the end points other than cardiovascular events or death (n = 15) or used RAAS inhibitors simultaneously in both trial arms (n = 7) or were not relevant (n = 40). Finally, 38 RCTs assessing the association of cardiovascular outcome or cardiovascular or total mortality with ACEIs or ARBs were included in the meta-analysis [2–14, 22–47]. As shown in Fig. 1, literature research process was summarized by a chart flow. Baseline characteristics of all selected studies are detailed in Table 1. A total of 47,662 subjects were included with a mean/median follow-up ranged from 12 weeks to 4.5 years. Of all 38 studies, six (n = 8404) trials compared ARBs with placebo [13, 43–47], while 32 trials (n = 39,254) compared ACEIs with various control therapies (13 arms (n = 10,134) compared ACEIs [2–6, 23–30] with placebo treatment; 10 arms (n = 8714) in which the comparator was active treatment [7, 8, 31–38]; and 9 arms (n = 20,406) compared ACEIs with ARBs [9–12, 14, 39–42]). Two independent investigators (Y. X. and D. X.) assessed the quality of the studies included. There were 32 studies of good quality (Jadad score ≥ 3) with low risk of bias and six studies of low quality (Jadad score < 3) with high risk of bias.

Effect of ACEIs and ARBs on all-cause mortality
Thirty-two studies [2–12, 14, 23–42] reported the effect of ACEIs on all-cause mortality in a total of 39,254 HF patients with moderate heterogeneity in overall analysis (I² = 44%, p = 0.005). ACEIs were associated with a statistically significant 11% reduction in all-cause mortality (RR: 0.89, 95% CI: 0.83–0.96, p = 0.001, Fig. 2). Similar findings were observed when ACEIs were compared with placebo treatment (p < 0.001, Fig. 2). There was no evidence of publication bias (p = 0.833).

Moreover, 15 studies [9–14, 39–47] reported the effect of ARBs on all-cause mortality in a total of 28,814 HF patients with no significant heterogeneity in overall analysis (I² = 26%, p = 0.17). ARBs were not associated with a reduction in all-cause mortality (RR: 1.03, 95% CI: 0.98–1.08, p = 0.28, Fig. 3). Similar findings were observed when comparing with placebo or ACEIs (p ≤ 0.60, Fig. 3). And there was no evidence of publication bias (p = 0.921).

Fig. 1 Flow diagram of selection strategy. ACEI, angiotensin-converting enzyme inhibitor, ARB, angiotensin II receptor blocker
| Study, year    | No of patients | Drugs | Follow-up, w | Cause of heart failure | Risk factors | Jadad Score |
|---------------|----------------|-------|--------------|------------------------|--------------|-------------|
|               |                |       | Cause | MI, % | HTN, % | TCM, % | NICM, % | VHD, % | DM, % | HTN, % | AF, % |          |
|               |                |       | Baseline characteristics | Men, % | Age, y | LVEF, % |          |          |          |          |          |          |          |
| ARBs vs Controls |                |       |       |        |        |        |          |          |          |          |          |          |          |
| Havranek [43], 1999 | 218            | Irbesartan | Placebo | 82 | 60 | ≤0.40 | 12 | – | – | 67 | – | – | – | – | – | 2 |
| STRETCH [45], 1999 | 844            | Candesartan | Placebo | 68 | 62 | 0.35–0.45 | 12 | – | 29 | 71 | 2 | 2 | – | – | – | 4 |
| SPICE [44], 2000 | 270            | Candesartan | Placebo | 69 | 66 | <0.35 | 12 | – | 4 | 71 | 16 | 11 | 48 | 34 | 24 | 5 |
| ARCH-J [47], 2003 | 292            | Candesartan | Placebo | 78 | 64 | ≤0.45 | 24 | 25 | 7 | – | 57 | 8 | – | – | – | 4 |
| Val-HeFT [46], 2001 | 5010           | Valsartan | Placebo | 80 | 63 | ≤0.40 | 100 | – | 7 | 57 | 31 | – | 26 | – | 12 | 5 |
| CHARM-Alternative [13], 2003 | 2028 | Candesartan | Placebo | 68 | 67 | ≤0.40 | 135 | – | 6 | 68 | – | – | 27 | 50 | 25 | 5 |
| ARBs vs ACEIs |                |       |       |        |        |        |          |          |          |          |          |          |          |          |
| REPLACE [42], 2001 | 378            | Telmisartan | Enalapril | 89 | 64 | ≤0.40 | 12 | – | – | – | – | – | – | – | – | 4 |
| HEAVEN [40], 2002 | 141            | Valsartan | Enalapril | 75 | 67 | ≤0.45 | 12 | – | – | – | 87 | – | – | – | – | 3 |
| Dickstein [39], 1995 | 166            | Losartan | Enalapril | 78 | 64 | <0.35 | 12 | – | – | 69 | 3 | 12 | – | 23 | – | 3 |
| ELITE [14], 1997 | 722            | Losartan | Captopril | 67 | 73 | ≤0.40 | 48 | – | – | 68 | – | – | 25 | 57 | 23 | 4 |
| ELITE II [10], 2000 | 3152           | Losartan | Captopril | 70 | 71 | ≤0.40 | 72 | – | – | – | – | – | 24 | 49 | 30 | 5 |
| RESOLVD [12], 1999 | 768            | Candesartan | Enalapril | 84 | 63 | <0.40 | 43 | – | – | 72 | – | – | – | – | – | 5 |
| OPTIMAAL [9], 2002 | 5477           | Losartan | Captopril | 71 | 67 | ≤0.35 | 130 | 100 | 0 | 0 | 0 | 17 | 36 | 10 | 5 |
| VALIANT [11], 2003 | 14703          | Valsartan | Captopril | 69 | 65 | ≤0.35–0.45 | 107 | 100 | 0 | 0 | 0 | 23 | 55 | – | 5 |
| Lang [41], 1997 | 116            | Losartan | Enalapril | 78 | 58 | ≤0.45 | 12 | – | 4 | 47 | 44 | 3 | – | – | – | 3 |
| ACEIs vs Controls |                |       |       |        |        |        |          |          |          |          |          |          |          |          |
| AIRE [2], 1993 | 1986           | Ramipril | Placebo | 74 | 65 | – | 60 | – | – | – | – | – | 12 | 28 | – | 5 |
| Balpitt [23], 1998 | 169            | Captopril | Placebo | – | – | – | 24 | – | – | – | – | – | – | – | – | 2 |
| CASSIS [24], 1995 | 96             | Enalapril | Placebo | 83 | 58 | <0.40 | 12 | – | – | 70 | 30 | – | 23 | – | – | 3 |
| Chalmers [25], 1987 | 130            | Lisinopril | Placebo | 69 | 58 | – | 12 | – | 13 | 48 | 30 | 8 | – | – | – | 2 |
| Cofer [26], 1992 | 172            | Benazepril | Placebo | – | – | ≤0.35 | 12 | – | – | – | – | – | – | – | – | 2 |
| CONSENSUS [3], 1987 | 253            | Enalapril | Placebo | 70 | 71 | – | 27 | – | – | 73 | 15 | 26 | 23 | 25 | 58 | 3 |
| FEST [27], 1995 | 308            | Fosinopril | Placebo | 74 | 63 | ≤0.35 | 12 | – | – | – | – | – | – | – | – | 4 |
| FHFSG [28], 1995 | 241            | Fosinopril | Placebo | 80 | 62 | ≤0.35 | 24 | – | – | – | – | – | – | – | – | 3 |
| Lechat [29], 1993 | 125            | Perindopril | Placebo | – | – | – | 12 | – | – | – | – | – | – | – | – | 3 |
| Newman [30], 1988 | 105            | Captopril | Placebo | – | – | – | 12 | – | – | – | – | – | – | – | – | 2 |
| SAVE [4], 1992 | 2231           | Captopril | Placebo | 82 | 59 | ≤0.40 | 144 | 100 | – | – | – | – | 21 | 43 | – | 5 |
| SOLVD [5], 1991 | 2569           | Enalapril | Placebo | 80 | 61 | ≤0.35 | 166 | – | – | – | – | – | 26 | 42 | 10 | 5 |
| TRACE [6], 1995 | 1749           | Trandolapril | Placebo | 72 | 68 | ≤0.35 | 96–200 | 100 | – | – | – | – | 14 | 23 | – | 5 |
| Aguilar [31], 1999 | 345            | Captopril | Digoxin | 68 | 63 | – | 216 | – | – | – | – | – | – | – | – | 3 |
| Study, year          | No. of patients | No. of patients | No. of patients | No. of patients | No. of patients | No. of patients |
|---------------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|
| CARMEN [32], 2008  | 381             | 14              | 32             | 17              | 4               |
| CIBIS III [33], 2011| 217             | 72              | 61             | 12              | 3               |
| Cowley [34], 1994 | 209             | 32             | 48             | –               | –               |
| Dohmen [35], 1997  | 266             | 66             | 24             | <0.40           | –               |
| Hy-C [36], 1992    | 104             | 50             | 24             | 24              | >40             |
| IMPRESS [37], 2007 | 573             | 34             | 32             | 4               | –               |
| Northridge [38], 1999 | 5770          | 31             | 24             | 3               | –               |
| OVERTURE [7], 2002 | 45              | 48             | 21             | 3               | –               |
| V-HeFT II [8], 1991| 804             | 56             | 24             | 3               | –               |

**Table 1** Study characteristics (Continued)

- **No. of patients**: This column indicates the number of patients included in each study.
- **No. of patients**: This column indicates the number of patients in the treatment group.
- **No. of patients**: This column indicates the number of patients in the control group.
- **Follow-up, w**: This column indicates the duration of the follow-up period in weeks.
- **Jadad Score**: This column indicates the Jadad score, a measure of the methodology of a randomized controlled trial.
- **Risk factors**: This column lists the risk factors associated with heart failure, including age, sex, and other factors.

**Notes**:
- Data was absent in the original article.
- No number, LVEF: left ventricular ejection fraction, MI: myocardial infarction, HTN: hypertension, DM: diabetes mellitus, AF: atrial fibrillation, ACEI: angiotensin-converting enzyme inhibitors, ARB: angiotensin II Receptor Blockers, ICM: ischemic cardiomyopathy, NICM: non-ischemic cardiomyopathy, VHD: valvular heart disease, n: mean.
Fig. 2 Forest plot of angiotensin-converting enzyme inhibitors (ACEIs) compared with controls on all-cause mortality. Boxes and solid lines indicate RR and 95% CI, respectively for each study, and the diamonds and their width indicate the pooled RR and the 95% CI, respectively. M-H indicates Mantel-Haenszel. ACEI, angiotensin-converting enzyme inhibitor, ARB, angiotensin II receptor blocker.
Figure 4 showed the relation between the network of RCTs.

Effect of ACEIs and ARBs on CV mortality
Seventeen studies [3–6, 8–11, 14, 24, 32, 35, 36, 38, 40–42] reported the effectiveness of ACEIs for CV mortality in a total of 28,302 HF patients with moderate heterogeneity in overall analysis ($I^2 = 51\%$, $p = 0.009$). ACEIs were associated with a statistically significant 14% reduction in CV mortality (RR: 0.86, 95% CI: 0.78–0.94, $p = 0.001$, Fig. 5). Similar findings were observed when ACEIs treatment was compared with placebo treatment ($p < 0.001$, Fig. 5). However, when ACEIs were compared with active treatment or ARBs, ACEIs did not significantly reduce CV mortality. There was no evidence of publication bias ($p = 0.967$). The SAVE [4], TRACE [6] and VALIANT [11] study were conducted in patients with HF after myocardial infarction. After exclusion of these three trials, heterogeneity among the trials was not significantly different ($I^2 = 34\%$, $p = 0.10$, RR, 0.85, 95% CI: 0.76–0.95, $p = 0.005$).

Moreover, 11 studies [9–11, 13, 14, 40–42, 45–47] reported the effectiveness of ARBs for CV mortality in a total of 27,991 HF patients with no significant heterogeneity in overall analysis ($I^2 = 40\%$, $p = 0.08$). ARBs were associated with no reduction in CV mortality (RR: 1.01, 95% CI: 0.92–1.12, $p = 0.78$, Additional file 1: Figure S1). Similar findings were observed when ARBs were compared with placebo or ACEIs ($p \leq 0.50$, Additional file 1: Figure S1). And there was no evidence of publication bias ($p = 1.000$).

Meta-regression
Meta-regression was conducted in different ages ($p = 0.97$), causes of HF ($p = 0.90$), left ventricular ejection fractions ($p = 0.09$), follow-up weeks ($p = 0.41$) to observe effects of ACEIs treatment on all-cause mortality. The findings remained unaltered in these subgroup analyses. But, univariate meta-regression of ACEIs treatment on all-cause mortality varied by the types of ACEIs ($p = 0.004$). Captopril treatment reduced all-cause mortality by 9% (RR: 0.91, 95% CI: 0.85–0.98, $p = 0.008$) in
HF patients as compared with control treatment. However, enalapril treatment did not reduce all-cause mortality in HF patients as compared with control treatment (RR: 0.93, 95% CI: 0.85–1.02, \(p = 0.13\), Fig. 6). Of all these studies, one study compared ramipril with placebo and two studies compared lisinopril with placebo/active drugs. The results were shown in Fig. 5.

**Discussion**

**Major findings**

In this meta-analysis, we combined clinical trial data from 38 studies, which included 47,662 HF patients to assess the efficacy of RAAS inhibition on mortality. Overall, ACEIs reduce all-cause mortality by 11% and the corresponding value for CV mortality by 14%. However, ARBs have no significant effect on all-cause and CV mortality in HF patients. In head-to-head analysis, ACEIs are not superior to ARBs on all-cause and CV mortality. Thus, this meta-analysis provides compelling evidence that ACEIs are the most effective first-line treatment for preventing all-cause and CV mortality in HF patients.

RAAS inhibition has long been identified as a preferred first-line treatment for heart failure. However, previous studies indicated that there were different outcomes between AECIs and ARBs for heart failure. As early as 1987, CONSENSUS study [3] was conducted to evaluate the efficiency of enalapril in patients with HF. Six-month mortality in the enalapril group was 26% compared with 44% in the placebo group, giving a relative risk-reduction of 40% \((p = 0.002)\) and at 1 year, these proportions were 36% and 52% \((p = 0.001)\). After that, several studies demonstrated that ACEIs reduced all-cause and CV mortality in HF patients, particularly after myocardial infarction [4, 6]. However, most of ARBs are not proved to be effective on these crucial outcomes in HF patients. For example, the ELITE II study [10] found that losartan is not superior to captopril, although it has been suggested that the dose of losartan (50 mg) tested is not adequate. And in the CHARM-Alternative trial [13], candesartan did not reduce all-cause mortality in HF patients, but reduced the risk of CV death or HF hospitalization by 23% \((p = 0.0004)\). The Val-Heft study [46] showed the same results. These may due to the negative effect of ARBs on heart failure, which could be mediated through a vasoconstrictor-induced increase in blood pressure or a direct effect on cardiac and vascular tissues. So, more related studies are expected to conducted in this area. Besides, in some recent meta-analysis, Vark et al. [15] and Cheng et al. [16] presented that in patients with hypertension and diabetes, treatment with an ACEI resulted in a significant further reduction in all-cause and CV mortality, whereas ARBs had no benefit on these outcomes. These results are in agreement with our meta-analysis.

**Pharmacological mechanism**

From a pharmacological viewpoint, ACEIs can reduce the negative effects caused by binding of angiotensin II and its receptor by inhibiting the conversion of angiotensin I to angiotensin II. In addition, by restraining the degradation of angiotensin (1–7) and promoting its combination with Mas receptor, ACEIs may have effect on dilating blood vessels, anti-inflammatory and anti-fibrosis. Moreover, ACEIs can also reduce degradation of bradykinin and promote its role in \(\beta_2\) receptor, which contributes to dilation of blood vessels, anti-proliferation, endothelial protection and other positive effects [48, 49]. In contrast, RAAS blockade with ARBs is achieved by inhibiting the binding of angiotensin II to the angiotensin II type one receptor, which is believed to
mediate the harmful cardiovascular effects of angiotensin II due to the elevated level of angiotensin II by compensatory mechanism. These different pharmacological mechanisms may explain that ACEIs may be superior to ARBs in reducing CV events. Therefore, an ACEI agent may be a superior ARB antagonist in hypertension and heart failure.

**Heterogeneity**

There was low to moderate heterogeneity of analysis on the effect of ACEIs on all-cause and CV mortality. Meta-regression, sensitivity and subgroup analysis were conducted to estimate the influence of each study. Firstly, no evidence shows that the observed effects varied by age, causes of HF, left ventricular ejection fractions and follow-up weeks by meta-regression. However, different types of ACEIs may influence the effect on all-cause mortality, which means that captopril may be superior to enalapril in reducing all-cause mortality in HF patients. Secondly, the SAVE [4], TRACE [6] and VALIANT [11] study were conducted in patients with heart failure after myocardial infarction. After exclusion of these three trials, heterogeneity among the trials exploring the effect of ACEIs on CV mortality was not significantly different ($I^2 = 34\%, p = 0.10$, RR, 0.85, 95% CI: 0.76–0.95, $p = 0.005$). So, the significant heterogeneity

| Study or Subgroup | ACEI | Control | Risk Ratio | M-H, Random, 95% Cl | Risk Ratio | M-H, Random, 95% Cl |
|-------------------|-----|---------|------------|----------------------|------------|----------------------|
| 1.2.1 Placebo     |     |         |            |                      |            |                      |
| CASSIS 1995       | 0   | 48      | 6 48       | 0.1%                 | 0.08 [0.00, 1.33] |                      |
| CONSENSUS 1987    | 44  | 127     | 64 126     | 6.4%                 | 0.68 [0.51, 0.92] |                      |
| SAVE 1992         | 188 | 1115    | 234 1116   | 11.1%                | 0.60 [0.68, 0.96] |                      |
| SOLVD 1991        | 399 | 1265    | 461 1264   | 14.5%                | 0.66 [0.78, 0.96] |                      |
| TRACE 1995        | 226 | 876     | 288 873    | 12.5%                | 0.78 [0.68, 0.91] |                      |
| Subtotal (95% CI) | 3451| 3447    | 44.7%      |                      | 0.81 [0.72, 0.89] |                      |
| Total events      | 857 | 1053    |            |                      |            |                      |
| Heterogeneity: $\tau^2 = 0.00$; $\chi^2 = 5.51$, df = 4 ($P = 0.24$); $I^2 = 27\%$ | Test for overall effect: $Z = 4.29$ ($P < 0.0001$) |                      |

| 1.2.2 Active      |     |         |            |                      |            |                      |
| CARMEN 2008       | 14  | 190     | 13 191     | 1.5%                 | 1.08 [0.52, 2.24] |                      |
| Dohmen 1997       | 3   | 136     | 1 130      | 0.2%                 | 2.87 [0.30, 27.22] |                      |
| Hy-C 1992         | 9   | 44      | 26 60      | 1.8%                 | 0.47 [0.25, 0.89] |                      |
| Northridge 1999   | 1   | 23      | 0 22       | 0.1%                 | 2.88 [0.12, 67.03] |                      |
| V-HeFT II 1991    | 112 | 403     | 137 401    | 9.5%                 | 0.81 [0.66, 1.00] |                      |
| Subtotal (95% CI) | 796 | 804     | 13.1%      |                      | 0.80 [0.57, 1.11] |                      |
| Total events      | 139 | 177     |            |                      |            |                      |
| Heterogeneity: $\tau^2 = 0.04$; $\chi^2 = 5.09$, df = 4 ($P = 0.28$); $I^2 = 21\%$ | Test for overall effect: $Z = 1.34$ ($P = 0.18$) |                      |

| 1.2.3 ARB         |     |         |            |                      |            |                      |
| ELITE II 2000     | 24  | 370     | 12 352     | 1.7%                 | 1.90 [0.97, 3.75] |                      |
| HEAVEN 2002       | 4   | 71      | 1 70       | 0.2%                 | 3.94 [0.45, 34.41] |                      |
| Lang 1997         | 0   | 38      | 4 78       | 0.1%                 | 0.23 [0.01, 4.06] |                      |
| OPTIMAAL 2002     | 363 | 2733    | 420 2744   | 13.4%                | 0.87 [0.76, 0.99] |                      |
| REPLACE 2000      | 2   | 77      | 4 301      | 0.3%                 | 1.96 [0.36, 10.47] |                      |
| VALIANT 2003      | 830 | 4909    | 827 4909   | 15.6%                | 1.00 [0.82, 1.20] |                      |
| Subtotal (95% CI) | 9772| 10032   | 42.3%      |                      | 0.95 [0.83, 1.10] |                      |
| Total events      | 1422| 1498    |            |                      |            |                      |
| Heterogeneity: $\tau^2 = 0.01$; $\chi^2 = 11.73$, df = 6 ($P = 0.07$); $I^2 = 49\%$ | Test for overall effect: $Z = 0.68$ ($P = 0.50$) |                      |

| Total (95% CI)    | 14019| 14283 | 100.0%     | 0.86 [0.78, 0.94] |                      |
| Total events      | 2418 | 2728  |            |                      |            |                      |
| Heterogeneity: $\tau^2 = 0.01$; $\chi^2 = 32.53$, df = 16 ($P = 0.009$); $I^2 = 51\%$ | Test for overall effect: $Z = 3.23$ ($P = 0.001$) |                      |
| Test for subarous differences: $\chi^2 = 3.83$, df = 2 ($P = 0.15$); $I^2 = 47.8\%$ |                      |                      |

**Fig. 5** Forest plot of angiotensin-converting enzyme inhibitors (ACEIs) compared with controls on cardiovascular mortality. Boxes and solid lines indicate RR and 95%CI, respectively for each study, and the diamonds and their width indicate the pooled RR and the 95% Cl, respectively. M-H indicates Mantel-Haenszel. ACEI, angiotensin-converting enzyme inhibitor, ARB, angiotensin II receptor blocker.
was attributable to the different control treatment. There was no evidence of publication bias ($p = 0.721$).

Study strengths and limitations
Strengths of the present study are no other than the large sample size with a mean/median follow-up ranging from 12 weeks to 4.5 years and a high representativeness.

It has been acknowledged that there are some limitations to this study. Firstly, this analysis used aggregate data as reported or calculated in published articles, rather than data of individual patients. Secondly, there were a great deal of variations between the studied populations. For example, causes of heart failure differed from each other. In addition, these trials used different ACEIs or ARBs at a different dosage. It is likely that different ACEIs and ARBs may have a total different effect on the cardiac mortality. Moreover, the present study is unable to address whether the efficacy may be varied in HF patients with different ethnic backgrounds.

Conclusions
In 47,662 subjects, our meta-analysis shows that ACEIs, but not ARBs reduce all-cause mortality and cardiovascular deaths in HF patients. Thus, ACEIs should be considered as first-line therapy to limit excess mortality and morbidity in this population.

Additional file

**Additional file 1: Figure S1.** Forest plot of angiotensin II receptor blocker inhibitors (ARBs) compared with controls on cardiovascular mortality. Boxes and solid lines indicate RR and 95% CI, respectively for each study, and the diamonds and their width indicate the pooled RR and the 95% CI, respectively. Trials to the left of the vertical line showed a reduction in risk with the experimental intervention; those to the right showed an increase in risk with the experimental intervention. ACEI, angiotensin-converting enzyme inhibitor, MI, myocardial infarction, HF, heart failure, LVEF, left ventricular ejection fraction

**Abbreviations**
ACEIs: Angiotensin-converting enzyme inhibitors; ARBs: Angiotensin II receptor blockers; CI: Confidence interval; CV: Cardiovascular; HF: Heart failure; LVEF: Left ventricular ejection fraction; RAAS: Rennin angiotensin aldosterone system; RCT: Randomized clinical trial; RR: Risk ratio

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**Availability of data and materials**
The datasets used and/or analysed during the current study are available from the corresponding author on reasonable request.

**Authors’ contributions**
[DX, HL, WC, YS, YZ, LZ, CT, TG, YZ, JL] conception and design of the research. [CT, TG, JZ, YS, WC] acquisition of data. [CT, TG, LZ, JL, WC, YX, YZ] analysis and interpretation for the data. [CT, TG, LZ, JL, YZ, YS, JL] statistical analysis. [DX, HL] obtaining funding and supervising the work. [CT, TG, JZ, WC, LZ, YZ, JL] drafting the manuscript. [DX, HL, WC, YS, TG, JL] critical
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