Sifting and winnowing: Analysis of farmer field data for soybean in the US North-Central region

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Abstract
Field trials are commonly used to estimate the effects of different factors on crop yields. In the present study, we followed an alternative approach to identify factors that explain field-to-field yield variation, which consisted of farmer survey data, a spatial framework, and multiple statistical procedures. This approach was used to identify management factors with strongest association with on-farm soybean yield variation in the US North Central (NC) region. Field survey data, including yield and management information, were collected over two crop growing seasons (2014 and 2015) from rainfed and irrigated soybean fields (total of 3568 field-year observations). Fields were grouped into technology extrapolation domains (TEDs) that accounted for soil and climate variation and 9 TEDs were selected based on the number of fields needed to detect yield differences due to management as determined using power analysis. Average yield ranged from 2.5 to 5 Mg ha\(^{-1}\) across TEDs, with field yield distributions in half of the domains having a distributional peak that was close to maximum yields. Conditional inference trees analysis was chosen among 26 statistical procedures as the approach that best combines ability to detect and rank factors (and their interactions) with greatest influence on on-farm yield and relatively easy interpretation of results. Survey data from ca. 150 fields in each of the nine TEDs allowed us to identify key management factors influencing yields for an agricultural area that includes ca. 7 million ha sown with soybean. In five of the nine TEDs, highest yields were observed in early-sown fields. Other factors explaining on-farm yield variation were maturity group, and in-season foliar fungicide and/or insecticide application, but, in some cases, their influence on yield depended upon sowing date and water regime. While the approach proposed here cannot establish cause-effect relationships conclusively, it can certainly provide a focus to replicated field experiments in relation to which management factors to investigate. We believe that future agronomic studies based on farmer survey data can greatly benefit from ex-ante identification of most important TEDs (relative to crop area and production) as well as determination of minimum number of farmer survey data that needs to be collected from each of them based on expected yield differences and variability. The approach is generic enough to be applied in other crop producing regions as long as farmer data and associated climate and soil databases are available.

Keywords
Soybean, Survey data, On-farm yield, Statistical technique, Conditional inference tree

Disciplines
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Sifting and winnowing: Analysis of farmer field data for soybean in the US North-Central region

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ABSTRACT

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1. Introduction

Average crop yields will need to increase substantially during the next 33 years to meet expected food demand increase while avoiding massive expansion of cropland area (Tilman et al., 2011; Alexandratos and Bruinsma, 2012; Grassini et al., 2013). This challenge can be achieved by increasing the rate at which best management practices are identified and adopted for a particular soil-climate context. Replicated field experiments are used in agricultural research to test new technologies and management practices. In these experiments, researchers selectively manipulate a production factor and, by comparing final yield against the yield of a “control” treatment, the magnitude of the yield response and its economic profitability are assessed. A limitation of this approach is that it often examines the effect of management practices at a small number of sites and years due to practical constraints (e.g., costs, logistics, etc.). Hence, extrapolation of their findings is typically confined to a narrow range of environments. Likewise, field experiments cannot test the effect of a large number of production factors (and their interactions) on yield due to the large number of plots that would be needed. And, finally, the management selected as “background” for these experiments (e.g., sowing date, tillage method) will also influence crop responses to a given technology or management. Given these limitations, it is relevant to search for alternative, cost-effective approaches that provide an indication of the management practices that perform best for a given climate-soil context.

Farmer survey data can be utilized as a cost-effective source of information to identify yield constraints and fine-tune management practices so that these yield limitations can be ameliorated or eliminated (e.g., Calvino and Sadas, 2002; Sadas et al., 2002; Lobell et al., 2005; Tittonell et al., 2008). An advantage of using farmer data is that it allows examination of opportunities for yield increase within the range of current management practices that are both cost-effective and logistically feasible in farmer fields. Another advantage of using farmer data is that, if surveyed fields are properly contextualized relative to their biophysical environment, it is possible to explore and quantify management × environment interactions (Rattalino Edreira et al., 2017). Such assessment would allow identification of suites of management practices that perform best for a given environment and provide a focus to traditional, costly field experiments so that they can target those management practices with the most likely impact on crop productivity and input-use efficiency.

Statistical analysis of farmer self-reported data poses challenges that need to be addressed to make meaningful and unbiased inferences. For example, in field experiments, different levels of a given management or input are assigned to experimental units. These experimental units are carefully selected based on their similarity, in order to avoid confounding factors influencing yield and to minimize the error variance. Each treatment level is applied to several experimental units (‘replicates’) to obtain an estimate of average yield and its variation. In contrast, farmer data do not follow an experimental design and lack random allocation of experimental units and replication. Variation in soil, weather, and management practices across fields results in minimal control over error variance. Several management practices (or inputs) may be applied simultaneously, leading to multicollinearity, making interpretation of results more challenging (Hastie et al., 2001). Additionally, it may be the case that a given management practice does not appear to be significantly associated with yield simply because that practice has already been widely adopted across fields (e.g., cultivars with herbicide-resistance traits). Despite all these limitations, farmer data have the potential to give an indication of the most important yield-limiting factors in a given region, which can, in turn, then be tested in more detailed field trials to experimentally confirm cause-effect relationships.

We argue here that proper analysis of farmer field data, when evaluating the influence of management factors on yield, requires: (i) a biophysical spatial framework to cluster fields into groups with relatively similar climate and soil, (ii) use of appropriate statistical methods that can handle the nuances associated with the structure of farmer survey data and to identify management interactions, and (iii) a deep agronomic knowledge and understanding of the cropping system context to interpret results and translate them into practical recommendations. Application of a spatial framework to identify causes of yield gaps has been addressed in a previous study (Rattalino Edreira et al., 2017). A major limitation of this previous study, as well as other studies looking into the causes of yield gaps (e.g., Mercau et al., 2001, Sadas et al., 2002; Grassini et al., 2011, 2015; Silva et al., 2016), is that the analysis was limited to a comparison of management practices between high- versus low-yield fields or regressions between yield and individual or multiple management practices for a given climate-soil domain, without an explicit attempt to rank the importance of each management practice based on its influence on yield and to identify interactions.

In the present study, we addressed the second requirement listed above, that is, the use of a proper statistical technique to identify and rank management factors (and their interactions) influencing soybean yield in farmer fields. We focused on soybean fields in the North Central US region, which accounts for ca. 85% of US soybean production and ca. 30% of global production (FAOSTAT, 2016; USDA-NASS, 2016). The objective of this study was to utilize self-reported farmer data and multiple statistical techniques, together with a spatial framework, to identify the management practices with greatest influence on rainfed and irrigated soybean yields across diverse climate and soil conditions.

2. Materials and methods

2.1. Database description

Soybean yield and management practices data were collected from 3568 fields sown with soybean in 2014 and 2015 across 10 states in the US NC region: Iowa (IA), Illinois (IL), Indiana (IN), Kansas (KS), Michigan (MI), Minnesota (MN), Ohio (OH), North Dakota (ND), Nebraska (NE), and Wisconsin (WI) (Fig. 1). Detailed description of the database is provided elsewhere (Rattalino Edreira et al., 2017). The majority of surveyed fields were non-irrigated, except in Nebraska, where there were both rainfed (34%) and irrigated fields (66%) located within the same region. Maize was the predominant prior crop (88% of total fields). Average regional yield represents ca. 22 (rainfed) and 13% (irrigated) of the estimated yield potential, indicating a relatively small (but still exploitable) room for increasing farmer yields through fine tuning of current management practices (Rattalino Edreira et al., 2017).

Farmers reported data on field location, average yield (adjusted to 13% moisture content), and management practices, including sowing date, seedling rate, row spacing, variety name, tillage method, drainage system, total irrigation amount (for irrigated crops), seed treatment, fertilizer inputs, lime, manure, and pesticides (Table 1). Farmers also reported incidence of other field adversities such as pests, diseases, weeds, iron deficiency chlorosis, hail, waterlogging, and frost. Data were subjected to quality control to remove erroneous entries. Likewise, fields subjected to unmanageable field adversities (e.g., hail, frost, flooding) leading to substantial yield losses were excluded from the analysis. To do this, fields reported as affected by any of the aforementioned adversities were grouped within regions with similar soil and climate (see Section 2.2), and we excluded those that fall below the 25th percentile of the yield data distribution within each region-year. To summarize, we excluded data from fields affected by unmanageable adversities and that fell below the 25th percentile of the yield distribution in each climate-soil domain; these data were excluded from all the statistical analyses, as well as tables and figures presented here. We did not exclude fields that suffered from drought, heat stress, temporary waterlogging, or disease, insect or weed pressure. After quality control, the database contained data from a total of 3216 fields sown with soybean in 2014 and 2015 (92% of total surveyed fields). Fields were
grouped into narrow (≈ 18 cm), intermediate (≈ 38 cm), and wide (≈ 76 cm) row spacing (Table 1). Fields were classified based upon tillage method as (i) conventional (chisel and disk), (ii) reduced (strip-till, ridge-till, cultivator), and (iii) no-till. Fields were classified depending upon seed treatment (ST) as untreated, fungicide-ST, and insecticide-ST. Because a substantial number of surveys did not indicate if ST included fungicide, or insecticide or both, we also used a generic ST class for the statistical analysis. Fields were also classified according to presence or absence of artificial drainage system such as new systematic tiles, old clay tiles, etc.

Mean pH was calculated for the topsoil (0–30 cm) and subsoil (30–150 cm) in each field from the SSURGO database (https://weboilsurvey.sc.egov.usda.gov/App/WebSoilSurvey.aspx) (Table 1). Mean pH for a given field was derived from the pixel pH distribution within each field (ca. 9 pixels per field). Using the mode instead of the mean would have resulted in a negligible change in the calculated field-level pH (< 1%). To account for differences in slope and terrain across a field, which could influence the crop water balance and final seed yield, we calculated the topography wetness index (TWI) for each field (Table 1). TWI has been used to characterize the potential for surface run-off and run-on in landscapes; hence, it can be used indirectly to assess the influence of field topography on crop productivity (Moore et al., 1993). High values are associated with flat terrain whereas smaller values are associated with more uneven fields (e.g., fields with slopes). TWI is usually correlated with other soil attributes, including soil organic matter, soil texture, and phosphorous content; hence, higher TWI values are generally associated with more productive soils. TWI was calculated using the rsaga.wetness.index package in R (R development Core team, 2016) using the 30-m resolution National Elevation Dataset (USDA:NRCS:Geospatial Data Gateway; https://datagateway.nrcs.usda.gov/). Mean TWI for a given field was derived from the pixel TWI distribution within each field.

### 2.2. Field clustering

Fields were aggregated in clusters based on their biophysical properties using a technology extrapolation domain (TED) spatial framework (Rattalino Edreira et al., 2017; http://www.yieldgap.org/web/guest/cz-ted). Briefly, TED framework delineates regions based on: (i) annual total growing degree-days (10 classes), (ii) aridity index (10 classes), (iii) annual temperature seasonality (3 classes), and (iv) plant-available water holding capacity in the rootable soil depth (10 classes; 50-mm class interval). Each TED corresponds to a specific combination of the four aforementioned parameters. For our analysis, we selected...
fields located within nine TEDs (Fig. 1). These TEDs included only rainfed fields (1R, 2R, 3R, 4R, 5R, 6R), only irrigated fields (8I and 9I), and both (7R-I). These TEDs portray well the range of climates and soils within the US NC region, including 7 million ha annually sown with soybean, which represent 21% of US soybean area. Detailed description of the field clustering by TEDs is provided elsewhere (Rattalino Edreira et al., 2017). Because water supply and management is very different between irrigated and rainfed fields, we treated 7R and 7I separately for the descriptive analysis. However, for the statistical analysis, we pooled the data from irrigated and rainfed fields to identify interactions between water regime and management practices. Because not all of the 3568 surveyed fields were located within one of the selected TEDs, our analysis used data from a subset of 1373 fields. There were more than 98 fields per TED, with an average of 153 fields per TED. This number of fields per TED represented a good compromise between maximizing the number of TEDs and having a reasonable number of fields per TED to detect yield differences due to management practices (see Sections 2.3 and 3.1).

2.3. Statistical analysis

2.3.1. Power analysis to establish number of fields required to detect yield differences

Ex-ante power analysis was used to determine the number of fields (sample size) needed to detect statistically significant yield differences due to management practices with a reasonable high level of confidence. Here, sample size was evaluated by power analysis based on realistic estimates of expected yield differences and yield variability for rainfed and irrigated soybean fields using SAS v.9.4 software (SAS Institute Inc., 2016). Different scenarios of expected yield difference between levels of binary variables and standard deviation of yield were explored. Our power analysis was constrained to agronomically-relevant ranges of seed yield variation and yield responses, which were chosen based on the range of yield variability across TEDs, measured with the standard deviation (SD), and yield differences due to management practices reported in the literature for soybean. For every yield difference and standard deviation combination, 500 random samples with different sample size (from 10 to 800 fields) and normally distributed data were created. Then, each sample was evaluated at 5% significance level using one-way analysis of variance. The power of each sample size for every combination of input parameters was the proportion of times that a given yield difference was detected at 5% significance level.

2.3.2. Statistical analysis to identify drivers of on-farm yield variation

The second part of the statistical analysis involved the use of multiple statistical procedures to identify the management and soil variables with the strongest influence on yield within each TED (Table 2). A total of 26 statistical procedures were used. These procedures utilize variable selection features and are commonly used in studies with unstructured data (e.g., observational studies) that contain multiple independent variables. Additionally, regression procedures, such as LASSO (least squared shrinkage operator) and elastic net, have desirable properties that can mitigate data multi-collinearity issues (Zou and Hastie, 2005; Dormann et al., 2013).

For each TED, 22 independent variables (Table 1) were ranked in descending order based on the frequency in which each variable was identified as statistically significant across the 26 statistical models. Hence, if a given variable was detected as statistically significant across all fitted models, the frequency would sum up to 26. Ranking the variables using a weighted sum based on their similarity (e.g., stepwise, backward elimination and forward selection) would have resulted in very similar ranks to those obtained with our simple frequency sum.

Regression trees analysis has been used in previous studies to identify yield constraints in farmer fields located within small geographic regions (e.g., Lobell et al., 2005; Tittonell et al., 2008; Ferraro et al., 2009). This method does not have assumptions relative to data distribution, with appealing features for survey data analysis, including automatic variable selection, interpretability of interactions between variables, and ability to handle missing data (Hastie et al., 2001). It can handle categorical and continuous explanatory variables without statistical distribution assumptions, it is robust in the presence of outliers, multicollinearity, and heteroskedasticity, and can reveal interactions among factors. Nevertheless, this approach has been criticized for lack of concept of statistical significance (Mingers, 1987), data overfitting, and selection bias towards covariates with many levels and many missing observations (Hothorn et al., 2006). Conditional inference trees, which is the focus of the third part of our analysis, have been proposed as an alternative to regression trees as the former overcome the bias and overfitting issues by utilizing the distributional properties of the data (Hothorn et al., 2006). This method estimates a relationship among several variables by binary recursive partitioning in a conditional inference framework using distributional properties of variables (Hothorn et al., 2006).

The conditional inference tree analysis was performed using the partykit package in R (R development Core team, 2016). Application of conditional inference trees to analyze combined data from multiple experiments has been described in Mourtzinis et al. (2018). Briefly, the algorithm tests the null hypothesis of independence between the response variable (i.e., yield) and any of the input variables (i.e., management and field variables; see Table 1). The algorithm selects the input variable with strongest association, measured by a p-value, with the response variable. Then, a binary split is implemented in the

| Method | Selection criteria |
|--------|---------------------|
| Stepwise | AIC-1000 model average-80% random sampling as previous with 20% partition for validation |
| Backward elimination | AIC-1000 model average-80% random sampling as previous with 20% partition for validation |
| Forward selection | AIC-1000 model average-80% random sampling as previous with 20% partition for validation |
| Least angle regression (LAR) | AIC-1000 model average-80% random sampling as previous with 20% partition for validation |
| Least squared shrinkage operator (LASSO) | AIC-1000 model average-80% random sampling as previous with 20% partition for validation |
| Group LASSO | AIC-1000 model average-80% random sampling as previous with 20% partition for validation |
| Adaptive LASSO | AIC-1000 model average-80% random sampling as previous with 20% partition for validation |
| Elastic net | AIC-1000 model average-80% random sampling as previous with 20% partition for validation |
| Random Forest regression | Number of trees = 1000, number of permutations = 1000 |

AIC: Akaike information criterion.
selected input variable (node) and all steps are recursively repeated. The terminal node accounts for the final subset of fields. The result of this procedure is a graph that looks like a tree. The sizes of intermediate and terminal nodes are defined according to pre-specified criteria. In this analysis, the criterion for the independence test was based on univariate p-values (alpha = 0.05). To ensure adequate power, besides the p-value, we ensured that each intermediate node account for a minimum of 33% of total observations, and a terminal node should contain a minimum of 11% (one third of observations in an intermediate node). All these criteria must be met at every step of the algorithm so that a variable can qualify for a split. To avoid overfitting and enhance interpretability, the maximum tree depth was set to 10 nodes. Explanatory power of the conditional inference tree was calculated with the coefficient of determination ($R^2$) and root mean square error (RMSE).

Sensitivity of the results due to the chosen criteria was assessed by repeating the analysis with different combinations of minimum number of fields per intermediate node (20–40% of total observations), and tree depth up to 20 nodes. To avoid overfitting and development of low-power models, terminal nodes were not allowed to contain less than 11% of total fields as this would result in nodes with low number of fields (< 5). In all models, regardless of the chosen criteria, the identified primary and secondary important variables and their thresholds were identical. Only the tertiary and least important variables varied in a few larger trees (> 10 nodes) and the goodness of fit of these expanded models was not substantially improved (< 5%).

To identify putative factors with greatest influence on on-farm soybean yield, and their interactions, all management practices reported by farmers, as well as pH and TWI values, were included in the conditional inference tree analysis. The analysis was performed separately for each TED. Data were pooled across years for the analysis because TEDs explained 31× more of the variation in farmer yields than year or TED × year interaction (Rattalino Edreira et al., 2017). TEDs accounted for variation in plant-available water holding capacity in the rootzone, soil depth, which is inherently correlated with other soil parameters such as soil texture, soil organic matter, and soil depth. In contrast, plant-available water holding capacity is not necessarily correlated with pH and TWI, which justifies inclusion of these two parameters as independent variables in the analysis. Furthermore, these two parameters are potentially manageable in a given field through pH correction and artificial drainage. Soybean varieties were described relative to their maturity group (MG), using the latter as one of the independent variables in our models. Maturity groups are usually designated using triple zero, double zero, zero and Roman numerals from I to X for very short- and long-season varieties, respectively. We did not attempt to quantify the influence of specific varieties on yield given the multitude of varieties (ca. 2000) sown across farmer fields (which will would leave us with very few observations per variety and per TED) and the rapid varietal turnover over time, which would make any inference about variety become obsolete very quickly.

3. Results

3.1. Required number of fields per TED to detect yield differences due to management

Sample size needed to reach power ≥ 0.8 for different expected yield differences and yield variability is shown in Fig. 2. High and low SD lines corresponded to hypothetical environments with respective high and low yield variation. For example, the SD for selected TEDs in our study ranged from 510 (irrigated and favorable rainfed environments) to 790 kg ha$^{-1}$ (rainfed environments), with an average of 600 kg ha$^{-1}$. The magnitude of the yield difference reflects the expected yield response to a management factor or applied input. For example, a previous study indicated that, on average, foliar fungicide and/or insecticide application increased yield by ca. 300 kg ha$^{-1}$, while a 4-week delay in sowing after end of April would reduce yield ca. 600 kg ha$^{-1}$ (Rattalino Edreira et al., 2017). As indicated previously, each of our selected TEDs included > 98 fields, with an average of 153 fields per TED. Such a sample size seems adequate to detect a significant yield difference with power ≥ 0.8 for most of the variables considered in the analysis that are expected to influence farmer yields, especially in TEDs with low yield variation (Fig. 2). In contrast, our analysis will have less power for testing yield differences on their statistical significance in environments with high yield variation. Fortunately, our selected TEDs corresponded to the first category of environments, as indicated by the small within-TED coefficients of variation for farmer yield (see Fig. 4).

3.2. Soybean management in the US NC region

Descriptive analysis for soybean management practices in each TED was summarized in Table 3 and Fig. 3. The study region was characterized by diversity of soil types, weather, and management practices.

Except for the alkaline subsoil in TEDs 3 and 9 (pH ≈ 8), average pH in the topsoil and subsoil ranged between 6 and 7.5, with the subsoil exhibiting slightly higher pH (Fig. 3 C-D). Higher TWI values in fields in TEDs 2, 7I, 8 and 9 indicated a smaller run-off potential and favorable soils compared to fields in other TEDs (Fig. 3H). Topsoil and subsoil pH and TWI varied greatly across fields within some of the TEDs (e.g., TEDs 1 and 6), which further justified their inclusion as independent variables in the statistical analysis. Average sowing date varied by up to 2 weeks among TEDs, from early-May to late-May in the southern (TED 2, 9) and northern (TED 3) regions, respectively (Fig. 3A). Most varieties sown in farmer fields belong to MGs 2 and 3, except for fields located in the north-west region (TED 3; MGs 0 and 1) (Fig. 3B). Narrow (≈ 18 cm) and intermediate (≈ 38 cm) row spacing prevailed across TEDs located in rainfed production environments; in contrast, wider row spacing (≈ 76 cm) was dominant in irrigated fields (Table 3). Seedling rates ranged from 35 to 45 seeds m$^{-2}$ (Fig. 3G), which, given a typical emergence rate of ca. 85–90% in soybean (Gaspar et al., 2017), indicate that seedling rates used by farmers are much higher than those required to achieve a plant density that maximize yield (27–32 plants m$^{-2}$; De Bruin and Pedersen, 2008). Higher seedling rates (ca. 10%) were observed in the eastern (TEDs 1 and 2) and western fringes (TEDs 3, 8 and 9) of the US NC region.

Applied fertilizer amounts (in fields that received fertilizer) ranged from 5 to 245 kg ha$^{-1}$ ($P_2O_5$) and from 10 to 340 kg ha$^{-1}$ ($K_2O$), respectively, with rates increasing following a west-east gradient (Fig. 3E–F). Starter N fertilizer (i.e., a small N fertilizer application at sowing) was rarely applied in fields located in the central and eastern...
parts of the US NC regions ( < 10% of fields) (Table 3). About 10–20% fields in the western fringe of the region received N starter (TEDs 7, 8, 9), with this frequency increasing up to ca. 40% in the TED located in the north-west region (TED 3). This TED also has the largest frequency of tilled fields (55%). In contrast, no-till was the most common tillage method across the rest of the TEDs. Frequency of fields with artificial drainage followed the east-west gradient in seasonal precipitation, increasing dramatically from < 30% fields with artificial drainage systems in the western fringe of the US NC region to > 70% fields with drainage systems in the central and eastern regions (Table 3). Harvest and/or grazing of the residue left by previous maize crop were rarely practiced, except for 35–50% of fields located in western TEDs (TEDs 7, 8, and 9). Lime and manure were applied in < 20% of fields across TEDs, with most of these fields located in the central and eastern regions (Table 3).

Use of a seed treatment, which usually includes fungicide and/or insecticide, was a widespread practice across all TEDs, with seed being treated in > 80% of fields (Table 3). The frequency of fields that received foliar fungicide and/or insecticide applications ranged from 20 to 50% across TEDs and number of fungicide- and insecticide-treated fields were similar, in part because farmers tended to apply fungicide and insecticide together. A notable exception was the north-west TED (TED 3) where frequency of fields only treated with insecticides was much higher in relation with fungicide-treated fields (40 versus 11%). On average, 15% of surveyed fields reported incidence of soybean cyst nematode (SCN, *Heterodera glycines Ichtinoche*); however, it was remarkable that ca. 35% of the farmers did not know (because of lack of soil testing) about the incidence of this pest in their soybean fields.

Examination of TED 7 allowed assessing differences in management practices between rainfed and irrigated fields within the same climate-soil context (Fig. 3, Table 3). For example, irrigated fields were sown (ca. 7 days) earlier and with earlier maturing varieties (0.5 Mg difference) than rainfed crops. Likewise, a greater frequency of irrigated fields were tilled, received seed treatment and foliar fungicide, and used wider row spacing relative to rainfed fields located within the same TEDs. Higher TWI in irrigated versus rainfed fields indicated that the former were located in positions of the landscape with smaller surface runoff potential.

### Table 3
Description of management practices across technology extrapolation domains.

| Production factor (% fields) | Technology Extrapolation Domains (TEDs) |
|-----------------------------|----------------------------------------|
|                             | 1R  | 2R  | 3R  | 4R  | 5R  | 6R  | 7R  | 7I  | 8I  | 9I  |
| Inputs                      |     |     |     |     |     |     |     |     |     |     |
| Seed treatment              | 83  | 95  | 95  | 92  | 89  | 92  | 86  | 98  | 90  | 81  |
| Foliar fungicide            | 20  | 38  | 11  | 40  | 47  | 39  | 20  | 24  | 20  | 19  |
| Foliar insecticide          | 19  | 36  | 40  | 43  | 40  | 24  | 18  | 19  | 16  | 18  |
| Starter N fertilizer        | 7   | 0   | 39  | 5   | 6   | 3   | 14  | 10  | 11  | 18  |
| Lime                        | 10  | 23  | 0   | 15  | 4   | 16  | 16  | 10  | 3   | 0   |
| Manure                      | 12  | 12  | 0   | 10  | 11  | 16  | 4   | 12  | 0   | 0   |
| Field & crop management     |     |     |     |     |     |     |     |     |     |     |
| Artificial                     | 69  | 73  | 36  | 88  | 88  | 83  | 20  | 12  | 4   | 18  |
| Residue management:          |     |     |     |     |     |     |     |     |     |     |
| Grazed                       | 0   | 1   | 1   | 1   | 0   | 7   | 22  | 20  | 24  | 34  |
| Harvested                    | 6   | 23  | 1   | 3   | 2   | 0   | 15  | 17  | 16  | 19  |
| Tillage method:              |     |     |     |     |     |     |     |     |     |     |
| No-till                      | 60  | 44  | 20  | 48  | 59  | 52  | 72  | 67  | 50  | 90  |
| Reduced till                 | 17  | 19  | 25  | 25  | 19  | 20  | 14  | 13  | 17  | 5   |
| Conventional till            | 23  | 37  | 55  | 27  | 22  | 28  | 14  | 20  | 33  | 5   |
| Row spacing:                 |     |     |     |     |     |     |     |     |     |     |
| Narrow (< 18 cm)             | 18  | 31  | 25  | 14  | 2   | 13  | 2   | 10  | 14  | 14  |
| Intermediate (< 38 cm)       | 60  | 61  | 49  | 35  | 64  | 47  | 53  | 29  | 22  | 22  |
| Wide (< 76 cm)               | 22  | 8   | 26  | 51  | 34  | 40  | 45  | 61  | 64  | 64  |
| Adversities                  |     |     |     |     |     |     |     |     |     |     |
| Iron chlorosis deficiency    | 26  | 0   | 20  | 28  | 2   | 25  | 0   | 0   | 1   | 4   |
| Soybean cyst nematode:       |     |     |     |     |     |     |     |     |     |     |
| Yes                          | 16  | 15  | 7   | 28  | 26  | 11  | 7   | 19  | 13  | 7   |
| Unknown                      | 38  | 50  | 41  | 38  | 34  | 62  | 40  | 22  | 19  | 10  |

### 3.3. Yield variation among and within TEDs

Average soybean yield ranged from ca. 2.5 Mg ha⁻¹ in short-season rainfed environments (TED 3) to ca. 5 Mg ha⁻¹ in favorable irrigated areas (TEDs 8 and 9) (Fig. 4). Field-to-field variation within TEDs (quantified using the coefficient of variation (CV)) decreased with increasing average TED yield ($R^2 = 0.75$, $P < 0.05$). Yield variability within TEDs (range: 11–23%) was similar to the yield variation among TEDs (average CV = 16%), indicating that substantial field-to-field variation in yield remained after farmer fields were clustered based upon their TEDs. The high within-TED yield variation reflected the influence of other factors not accounted by the TEDs such as management practices, pH, and TWI.

For all TEDs, the distribution of soybean farmer yields was negatively skewed, although this pattern was more evident (skewness < −0.10) in half of the domains (Fig. 4). In other words, soybean field yields tended to have a distributional peak that was close to maximum yields, with skewing attributable to some low-yielding fields. Negatively skewed yield distributions have also been reported for other high-yield crop systems such as irrigated wheat grown in good soils in Yaqui Valley, northwestern Mexico (Lobell et al., 2005) and irrigated maize and soybean in Nebraska, USA (Grassini et al., 2011, 2014a). The degree of skewness was negatively correlated with the average TED yield ($R^2 = 0.35$, $P = 0.1$). To summarize, high-yield production environments exhibited smaller field-to-field yield variation, with a higher proportion of fields closer to maximum values compared with environments with lower yield.

### 3.4. Identification of candidate management factors with strongest influence on yield

Analysis of the soybean data with 26 different models helped us identify and rank the most important variables influencing soybean yield in each TED (Table 4). For example, row spacing, use of seed treatment, sowing date, topsoil pH, and TWI were the top-five ranked variables in TED 1. There were variables that consistently explained yield variation in a large number of TEDs. For example, sowing date
ranked amongst the top-fifteen variables in all TEDs and within the top-five variables in 6 of the 9 TEDs. Other variables that appeared as significant in most TEDs included the use of foliar fungicide and/or insecticide, seed treatment, TWI and pH, which appeared listed within the top-five variables in, at least, 4 of the 9 TEDs. Although there were similarities in the ranking among TEDs, there were also many
differences. For example, sowing date ranked first in TED 4R, but the same variable ranked last in TED 3R. Such an abrupt difference in rank is consistent with the remarkable difference in yield response to sowing date reported for these same TEDs (-33 versus -1 kg ha⁻¹ d⁻¹, respectively) by Rattalino Edreira et al. (2017), which was attributed to differences in water balance during the pod setting (R3-R5) phase.

Although the models in Table 2 identified and ranked variables in terms of importance (see Table 3), it is difficult to reveal and quantify interactive effects of 2- and 3-way interactions of continuous variables and, perhaps more importantly, it is difficult to interpret these ranks. Additionally, because in unstructured datasets not all levels of a variable always exist for all levels of the interacting variables, the risk of extrapolation beyond the actual range increases and, thus, interpretation of the interactions can be misleading. Amongst all statistical methods evaluated here, conditional inference tree analysis appeared as the most robust approach to identify and rank factors (and their interactions) with greatest impact on farmer yields while facilitating the interpretation of the results. Another advantage of this method, in relation with other statistical techniques, is that interactions between management and/or soil factors were restricted within the actual range of management and/or soil properties. For example, while we are aware about the power of random forest regression to develop yield prediction models, this method resembled a “black box” approach because interpretation of model results was extremely difficult.

The conditional inference tree analysis performed for rainfed fields located within one of the eastern TEDs (TED 1) is shown in Fig. 5. Sowing date was the most important variable influencing farmer soybean yields. On average, fields that were sown between day of year (DOY) 119 and 123 (late April and early May) yielded 4 Mg ha⁻¹ (left terminal node), which is 9% higher than average yield in late-sown fields. In late-sown fields (DOY from 124 to 167, which corresponded to late May-early June), highest yields were achieved in fields with relatively higher TWI (> 9.2) and lower subsoil pH (< 7.2), but these yields were still lower than those reported for early-sown fields. The three variables of the explanatory model (sowing date, TWI, and subsoil pH) captured approximately one third of total yield variability within the TED (R² = 0.29).

Sowing date was also the most important factor influencing soybean yields in TEDs 4R, 5R, 6R, and 8I (Table 5, Fig. 6). Remarkably, late-sown fields could not achieve yields comparable to early-sown fields.
under any suite of management practices and soil and terrain parameters. Foliar fungicide or insecticide was also identified as management factors increasing soybean yield in 5 of 9 TEDs (Fig. 7, Table 5). Higher yields were also generally related to high TWI, which may reflect a more favorable position in the landscape in relation to crop water supply and likely better soil quality (see Section 3.2). Other management factors influencing yield in at least one TED were row spacing, maturity group, tillage method, and seeding rate (Figs. 6 and 7, and Table 5).

Conditional inference trees also allowed us to capture M × E interactions. For example, MG was a significant secondary (TEDs 3 and 7) and tertiary key management practice (TEDs 3 and 9). In the short-season environment of TED 3, higher yields were associated with late MGs (Table 5). This finding was not biased by the latitudinal distribution of MG varieties within TED 3 as the influence of MG persisted even when the analysis was conducted separately for the southern and northern portions of this TED. In contrast, in favorable irrigated environments (TEDs 7 and 9), higher yields were achieved with early MGs (Fig. 7, Table 5). These findings are consistent with Specht et al. (1986, 2001), who noted that Midwestern U.S. full-season maturity cultivars in rainfed environments usually yield better than earlier-maturing ones, but generally yield less under irrigation. Drought can shorten reproductive development in the early-maturing cultivars aligning those stages with the hotter part of the growing season, which tends to exacerbate the impact of water deficit. Our analysis also revealed an interesting interaction between presence of soybean cyst nematode (SCN) and tillage: SCN led to lower yields in TED 6 (Fig. 6), but this yield reduction was 6% higher in no-till versus tilled fields.

Table 5
Summary of conditional inference trees in technology extrapolation domains (TEDs) 2R, 3R, 4R, 5R, 8I, and 9I. Values in brackets indicate number of fields (n) and average yield (Y, Mg ha⁻¹).

| TED# | N1                                      | N2                                      | N3                                      | N4                                      | [n, Y]          | R²  | RMSE (Mg ha⁻¹) |
|------|-----------------------------------------|-----------------------------------------|-----------------------------------------|-----------------------------------------|----------------|-----|----------------|
| 2R   | Row spacing (narrow)                    | Row spacing (intermediate, wide)        |                                        |                                        | [36,3.7]       | 0.10| 0.6            |
| 3R   | Foliar insecticide (yes)                | TWI (9.7–11.7)                         | TWI (8.2–9.7)                          | MG (0.9–1.5)                            | [82,4.3]       | 0.19| 0.48           |
|     | Foliar insecticide (no)                 | MG (0.08–0.9)                          | MG (0.08–0.6)                          | MG (0.6–0.9)                            | [23,2.9]       | 0.19| 0.48           |
| 4R   | Sowing date (DOY 143–175 DOY)           | Foliar fungicide (yes)                  |                                        |                                        | [23,3.9]       | 0.19| 0.48           |
|     | Sowing date (DOY 107–132)               | Foliar fungicide (no)                   |                                        |                                        | [27,3.7]       | 0.19| 0.48           |
|     | Subsoil pH (5.5–6.5)                    |                                        |                                        |                                        | [23,3.9]       | 0.19| 0.48           |
| 5R   | Sowing date (DOY 137–164)               | Foliar fungicide (no)                   |                                        | Sowing date (DOY 133–140)               | [27,3.7]       | 0.19| 0.48           |
|     | Subsoil pH (6.6–8.1)                    |                                        |                                        | Sowing date (DOY 141–161)               | [23,3.9]       | 0.19| 0.48           |
| 8I   | Sowing date (DOY 113–142)               | Foliar fungicide (yes)                  |                                        | Foliar fungicide (no)                   | [22,5.2]       | 0.26| 0.44           |
|     | Foliar fungicide (no)                   |                                        |                                        | Sowing date (DOY 113–124)               | [18,4.5]       | 0.26| 0.44           |
|     | Sowing date (DOY 125–142)               |                                        |                                        | Sowing date (DOY 125–142)               | [38,4.8]       | 0.26| 0.44           |
|     | Sowing date (DOY 107–132)               |                                        |                                        | Sowing date (DOY 107–132)               | [50,5.0]       | 0.26| 0.44           |
| 9I   | Seeding rate (30–36 m⁻²)                | TWI (8.4–9.1)                           | TWI (9.1–11)                           | MG (2.4–2.7)                            | [15,4.3]       | 0.34| 0.52           |
|     | Seeding rate (36–53 m⁻²)                | TWI (9.1–11)                           | MG (2.7–4.2)                           |                                     | [18,4.5]       | 0.34| 0.52           |

Nth: node number; TWI: topography wetness index; MG: maturity group; DOY: day of year.

Fig. 5. Conditional inference tree for technology extrapolation domain (TED) 1R, located in the eastern region of the US North-Central region. In each boxplot, the central rectangle spans the first to the third yield quartiles. The solid line inside the rectangle shows the mean, which is also reported in the bottom right corner. TWI = topography wetness index.
4. Discussion

Analysis of farmer survey data using multiple statistical methods and a spatial framework allowed us to identify the most critical management factors explaining field-to-field yield variation in major US soybean producing areas. It was remarkable that a reasonable number of TEDs (9) and number of fields per TED (ca. 150) was sufficient to identify the most important yield-limiting factors for an agricultural area that includes 7 million ha sown with soybean, which, in turn, represent 21% of US soybean area. A larger number of fields would have been needed for harsher rainfed environments with very high yield variation. We note, however, that these environments typically account for a smaller share of regional and national crop production. Overall, we believe that future agronomic studies based on farmer survey data can greatly benefit from ex-ante identification of most important TEDs in relation to crop area and production as well as determination of the minimum number of farmer survey data that needs to be collected from each of them based on expected yield differences and variability. Such an ex-ante analysis based on a spatial framework that accounts for key biophysical variables explaining yield variation and response to management practices, together with a power analysis to determine the minimum number of fields, can help prioritize resources by targeting TEDs with largest area and/or production (or other criteria) and ensuring that a sufficient (but not excessive number) of surveys are...
collected for each TED.

We used 26 different statistical models to analyze the data and we identified management and soil variables that were consistently correlated with soybean yield. For a given TED, selected variables varied across methods, which we attributed to the specific properties of each technique. For example, regression procedures such as LASSO, LAR, and elastic net have properties that can mitigate data multi-collinearity issues (Zou and Hastie, 2005; Dormann et al., 2013). Hence, as also noted in Krupnik et al. (2015), we found conditional inference tree analysis to outperform other statistical approaches when the goal is to both analyze and interpret unstructured farmer survey data. In our analysis for soybean in the US North-Central region, across TEDs, conditional inference models explained 10–44% of field-to-field yield variation using only one to four explanatory variables. Allowing the development of larger trees using increasingly smaller groups of fields would have inflated R² values at expense of higher uncertainty and increasing difficult to interpret the results. Likewise, our analysis was not intended to capture all possible sources of variability (e.g., climate variables). Instead, our objective was to identify key management, soil, and terrain factors influencing yield within each TED that could eventually be manipulated by farmers.

Sowing date and foliar fungicide and/or insecticide were the most persistent factors associated with yield variation. These results are consistent with findings from previous research based on farmer data collected from small geographic areas (e.g., Grassini et al., 2015), multi-year, multi-location replicated field experiments (e.g., Bastidas et al., 2008; Rowntree et al., 2013; Mourtzinis et al., 2016) and simulation modeling (e.g., Specht et al., 2014). But, in contrast to these previous studies, our analysis also exposed interesting interactions between management practices, for example, MG x water regime and nematodes x tillage, which are also consistent with experimental data (Specht et al., 1986, 2001; Conley et al., 2011). Interestingly, we could not detect a positive influence of narrow or intermediate row spacing on soybean yield despite the yield benefits of narrow row spacing reported in previous studies (Anaele and Bishnoi, 1992; Oplinger and Philbrook, 1992; Hanna et al., 2008; Chauhan and Opena, 2013). These contrasting results derived from on-farm data versus controlled experiments deserve further investigation. Sowing date exhibited a consistent association with yields, with diminishing yield as sowing date was delayed. It was remarkable that the yield loss due to late sowing could not be fully compensated by any combination of other management practices, such as seeding rate or row spacing. In other words, sowing date appears to play a major role in setting the yield potential for a given field, as other factors cannot compensate for late sowing. Hence, timely sowing appears as a key factor to increase the current soybean yields in the US NC region.

Identification of the causes for yield variation is needed but not sufficient for increasing farmer yields. For example, we identified sowing date as a key management factor explaining yield variation within the same TED. Hence, one would tend to think that it is relatively easy for a large number of farmers in the US North Central region to increase current soybean yield by sowing earlier, especially considering that early sowing date per se does not involve higher costs and labor. However, there are many reasons why farmers may still be reluctant to sow soybean earlier. The first constraint is a combination of farm logistics and cultural preference as many farmers only have one planter and they prefer to use it for sowing maize first. The second limitation is associated with biophysical factors (i.e., water excess, cold weather) that could delay sowing time in many years. Finally, farmers tend to overestimate the risk associated with seed chilling injury, early frost, and seed and/or plant stand loss associated with early sowing despite the well-documented benefits of early sowing and associated measures to reduce risk, for example, by using seed treatments or monitoring of soil temperature (e.g., Bastidas et al., 2008; Rowntree et al., 2013; Tenorio et al., 2016). Additionally, the current crop insurance program sets a limit to very early sowing for a given area. We note, however, that our analysis showed that a large number of the farmers are sowing soybean much earlier than other farmers within the same TED suggesting that closing the portion of the yield gap due to sowing date is possible through fine tune adjustment of farm logistics and a correct assessment (and mitigation) of risk level. Indeed, over the past three decades, farmers have persistently shifted average soybean sowing times in the US North Central region to earlier calendar dates at a rate of ca. 0.5 d year⁻¹ (Specht et al., 2014). The present study indicates that there is still large room for improving soybean yields by increasing the rate at which farmers shift toward early sowing.

In a broader context, given the growing pressure for increasing food production on existing cropland area, the approach used here represents a tremendous opportunity to help accelerate rates of yield gain and better prioritize research and extension programs in major crop producing regions of the world. Another strength of the approach is that it screens for suites of ‘best’ management practices within the context of the current cropping system; hence, it is able to capture the continuous changes in management practices as a result of farmer innovation and adoption of new technologies and identify emerging problems (Loomis, 1984; Passioua, 2010; Grassini et al., 2014b). While replicated field trials will still be needed to establish cause-effect relationships, the information derived from analysis on farmer data as presented here can provide a focus to these trials in regard to which factors (and interactions) to investigate. In other words, our approach can be considered as a complement to research based on randomized replicated field experiments. The approach proposed here is cost-effective and generic enough to be applied in any cropping system in the world as long as underpinning soil and climate data needed to contextualize farmer fields are available.

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