Community-based approaches for neonatal survival

Meta-analyses of randomized trial data

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Abstract:
Objective: To analyse the impact of community approaches to improving newborn health and survival in low-resource countries.

Methods: We updated previous meta-analyses of published cluster randomized trials of community-based interventions for neonatal survival. For each study we extracted baseline data on the context: geographical area; available facilities and staffing; immediate breastfeeding and facility births; and neonatal mortality. We also extracted data on the primary outcome (neonatal survival) and intermediate outcomes of the interventions (changes in immediate breastfeeding and facility births). We used forest plots and pooled sub-group analysis to seek patterns in associations between the effect size and the context or type of intervention (home-based counselling or women’s groups).

Findings: We included 17 trials, spanning years from 2001 to 2013. A 25 per cent reduction in neonatal mortality (relative risk, RR: 0.75; 95 per cent confidence interval, CI: 0.69–0.80) was found when pooling six studies in settings with 44 or more deaths per 1,000 live births. In lower-mortality settings (pooling six studies with 32 or fewer deaths per 1,000 live births) there was no evidence of an effect. We observed some evidence that community approaches had a stronger effect in south Asia than in sub-Saharan Africa. Community approaches had a lower impact on neonatal mortality in settings where at least 44 per cent of women delivered in a facility.

Conclusion: As neonatal mortality declined, the impact of community approaches on survival appeared to be lower, and the role of these approaches in supporting newborn care in weak health systems may need to be re-examined.

Keywords: community approaches, newborn health, neonatal mortality, neonatal survival, sub-Saharan Africa

Tables and figures: at the end of the paper.

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Introduction

Despite progress in reducing child deaths in the past 25 years, an estimated 2.6 million neonatal deaths occurred globally in 2015.\(^1\) Sustainable development goal (SDG) 3 included the target of no more than 12 deaths per 1000 live births in the first 28 days of life.\(^2\) To reach the target, more effective ways of delivering quality preventive and curative care need to be identified and monitored.

Approaches based on health promotion and on community empowerment and participation have long been promoted as part of formal health-care systems in low- and middle-income countries.\(^3,4\) Trials to improve maternal and newborn health through community approaches have focused on two approaches: (i) home-based counselling\(^5\) and (ii) participatory women’s groups.\(^6\) Both approaches promote appropriate care-seeking as well as improved home practices in newborn care. Home-based counselling focuses on health education and behaviour change to improve newborn care practices by mothers, such as immediate breastfeeding, dry cord care and appropriate health care (e.g. delivering in a health-care facility and seeking care for sick newborns). Women’s groups use an empowerment and problem-solving approach aiming similarly to improve care practices and care-seeking by mothers of newborns. The mechanisms of the effect of the home-based counselling strategies are backed by an analysis using the Lives Saved tool.\(^7\)

Previous meta-analyses have reported moderate effects on neonatal mortality of both home-based counselling and women’s groups. A meta-analysis of five proof-of-principle trials of home-based counselling in south Asia in 2010 found an almost 40% reduced risk of neonatal death (relative risk, RR: 0.62; 95% confidence interval, CI: 0.44–0.87).\(^5\) In response, the World Health Organization (WHO) recommended home visits to improve neonatal health in high neonatal mortality settings.\(^8\) However, trials of home-based counselling conducted in a larger population and in programme settings \(^9,10\) showed a smaller risk reduction for neonatal mortality (RR: 0.93; 95% CI: 0.85–1.01).\(^9\) A review of seven trials of women’s groups based on participatory learning and action cycles published in 2013 reported a 20% reduction in neonatal mortality (RR: 0.77; 95% CI: 0.65–0.90).\(^6\) The evidence prompted WHO to recommend community mobilization with women’s groups to improve maternal and neonatal health.\(^11\)

Factors reported to have the greatest impact on neonatal mortality include how successfully the intervention was implemented, as reflected by the proportion of pregnant
women participating in women’s groups; the inclusion of injectable antibiotics for treatment of possible severe bacterial infection; and home management of asphyxia. However, it is not clear how the women’s group approach works, or what is the interaction between community approaches and contextual factors, such as the characteristics of the health-care system.

In this paper we updated previous searches and meta-analyses of trials of home-based counselling and women’s groups in low-resource countries. The aim was to generate and test hypotheses about which factors may lead to weaker or stronger effects on neonatal survival. We examined associations between reductions in neonatal mortality and the context in which the trial took place or the characteristics of the local health system. We also assessed associations between reductions in mortality and the characteristics of the implementation.

Methods

Inclusion criteria and search methods

We reviewed cluster randomized trials evaluating community approaches to enhancing neonatal survival in low- and middle-income countries in April 2016, covering all studies published to this date. All trials compared neonatal mortality in pregnant women receiving the intervention with those receiving the local standard care (Table 1). We included trials of both home-based counselling and facilitated women’s groups delivered during pregnancy. Our starting point was two previously published reviews of five trials of home-based counselling interventions, and another five published between 2010 and 2013. One trial was excluded from the review as it was only quasi-experimental. We also included a review published in 2013 covering seven trials of women’s groups. To identify the most recently published trials we conducted a literature search of the PubMed and Web of Science online databases using the following search string ((("newborn" OR "neonatal" OR "maternal") AND mortality) AND trial), and identified trials of home-based counselling or women’s group interventions published between January 2013 and May 2016 in low- or middle-income countries (Fig. 1). We screened 1481 titles and identified one additional cluster randomized trial that examined women’s groups in rural eastern India. Another identified trial lacked a randomized design and was not included. Although they had been included in earlier meta-analyses by other authors, we excluded two non-randomized trials from our meta-analysis after an assessment of the risk of bias.
**Data processing**

Two authors independently assessed the risk of bias (allocation concealment, and method of data collection for neonatal mortality data) for each study included in the review using the Cochrane Collaboration tool.30

For each trial we extracted data on the study context (geographical area; baseline neonatal mortality rate; baseline proportion of births with infant breastfed immediately after delivery; baseline proportion of births in a facility); health system characteristics in the trial area (number of nurses and midwives per 1000 population; number of health facilities per 100 000 population); and type of intervention (home-based counselling or women’s groups). We also extracted data on the strength of the implementation (proportion of pregnant women visited in home-based counselling or attending women’s groups). Not all the variables were reported in all trials. The data were obtained from the published papers and through contacting authors. One author extracted data, which were subsequently checked by another author. We performed all analysis in Stata, version 13.0 (Stata Corp, College Station, United States of America).

The primary outcome for all studies was neonatal mortality. We also used immediate breastfeeding and facility births as tracer indicators for good newborn care practices. We calculated the changes in the proportions of women breastfeeding immediately after delivery and giving birth in a facility between baseline and endline separately for intervention and comparison groups (the difference-in-differences). When baseline figures were not available, we calculated the differences between the intervention and control groups at endline.

**Analysis**

We used the *metan* command in STATA to compute forest plots calculating the RR for neonatal mortality for each study based on the number of deaths and births reported in intervention and comparison groups at the end of the trial period. Heterogeneity was assessed and $I^2$ and $P$-values were tabulated together with the summary estimates to provide measures of heterogeneity. We used the forest plots to examine patterns in the effect size on neonatal mortality according to the study context or health system characteristics in the trial area. We also investigated patterns in neonatal survival according to features of the implementation. For the analysis of associations between intervention characteristics and the effect size we chose equal-sized groups. For example, we categorized the 17 trials into three groups based on neonatal mortality rates in the trial area: very high mortality settings ($\geq 44$ deaths per 1000 live
births), high mortality settings (33–43 deaths per 1000 live births) and moderately high mortality settings (≤ 32 deaths per 1000 live births).

**Results**

**Included trials**

We included 15 articles\textsuperscript{9,10,13–25} reporting 17 trials (two papers\textsuperscript{13,22} reported two studies each). Eight studies reported interventions using home-based counselling and nine were interventions based on women’s groups.

The trials took place in sub-Saharan Africa (Ghana, 1 trial; Malawi, 3 trials; United Republic of Tanzania, 1 trial) and in south Asia (Bangladesh, 4 trials; India, 6 trials; Nepal, 1 trial; Pakistan, 1 trial). All the trials were done in poor rural societies, except for the trial in an urban slum in India\textsuperscript{21} (Table 1).

All packages aimed to improve home-based newborn care by mothers, such as immediate and exclusive breastfeeding, thermal care, and safe and dry cord care; the home care arm from one study\textsuperscript{13} encouraged home treatment with antibiotics if referral was not possible (Table 1). Most trials reported coverage of these newborn practices as intermediate outcomes. Home-based behaviour change counselling involved visits to pregnant women at home by a community health worker or volunteer and sometimes also included community meetings. Women’s participatory groups took place in the community and were facilitated by trained community members who used problem-solving methods, such as action cycles. Both approaches included education and behaviour change communication to overcome challenges in health-care seeking and home newborn care practices.

All trials reported neonatal mortality as the main outcome, defined as the number of deaths in the first 28 days of life per 1000 live births in both sexes. Neonatal mortality data were obtained either from surveys or continuous surveillance in the target population. The trials were done in diverse contexts where the neonatal mortality rate ranged from 11 deaths per 1000 live births in an Indian urban slum\textsuperscript{21} to 84 deaths per 1000 live births in India.\textsuperscript{14} While the reported trials from Asia were from a period spanning the years 2001 to 2012, the reported trials from sub-Saharan Africa were from the years 2004 to 2013 (Table 1).

**Context characteristics**

We observed the largest reduction of neonatal mortality in settings with very high neonatal mortality. We calculated a 25% reduction in neonatal mortality (RR: 0.75; 95% CI: 0.69–0.80)
when pooling six studies \((P = 0.002\) for heterogeneity\) which took place in very high mortality settings of \(\geq 44\) deaths per 1000 live births. The effect on neonatal mortality was smaller \((RR: 0.89; 95\% CI: 0.83–0.95)\) when pooling five trials \((P = 0.392\) for heterogeneity\) in areas with high neonatal mortality of \(33–43\) deaths per 1000 live births, while there was no evidence of an effect on neonatal mortality \((RR: 0.94; 95\% CI: 0.88–1.01)\) when pooling six trials \((P < 0.001\) for heterogeneity\) in settings with moderately high neonatal mortality of \(\leq 32\) deaths per 1000 live births (Fig. 2; Table 2). The pattern of the largest reductions in settings with the highest neonatal mortality was observed for both home-based counselling and women’s group approaches (Fig. 3).

The pooled analysis suggested that the effects of the community approaches on neonatal mortality were stronger in the 12 pooled studies in south Asia \((RR: 0.82; 95\% CI: 0.78–0.86; P < 0.001\) for heterogeneity\), while there was no evidence of an effect in five studies in sub-Saharan Africa \((RR: 0.95; 95\% CI: 0.88–1.02; P = 0.193\) for heterogeneity\). None of the African studies, however, were done in a setting with very high neonatal mortality (Fig. 4; Table 2).

Overall, we did not observe any clear pattern of effects of immediate breastfeeding at baseline on neonatal mortality (Fig. 5; Table 2). However, trials done in settings with very high baseline neonatal mortality had lower rates of immediate breastfeeding (Fig. 5) and of facility births at baseline (Fig. 6). The mean baseline level of immediate breastfeeding was 31\% \((\text{range } 3–77\%)\) in very high mortality settings, 39\% \((\text{range } 11–54\%)\) in high mortality settings and 52\% \((\text{range } 19–82\%)\) in moderately high mortality settings. A similar trend was seen for facility births, whereby levels were 22\% \((\text{range } 3–48\%)\), 26\% \((\text{range } 2–52\%)\) and 43\% \((\text{range } 12–87\%)\) in very high, high and moderately high neonatal mortality settings, respectively.

The effect size of the community approaches was somewhat higher \((RR: 0.77; 95\% CI: 0.71–0.85)\) in pooled data from five studies \((P = 0.001\) for heterogeneity\) where the baseline level of facility births was low \((\leq 10\%)\). The effect was lower when pooling six studies with 11–43\% births in a facility \((RR: 0.85; 95\% CI: 0.80–0.91; P < 0.001\) for heterogeneity\) and five studies with \(\geq 44\%\) facility births \((RR: 0.90; 95\% CI: 0.83–0.97; P < 0.001\) for heterogeneity\); Fig. 6; Table 2).

Only nine trials reported the health-system characteristics of facilities in the trial area. The meta-analysis suggested a lower effect of the community-based approaches on neonatal mortality in settings with more health facilities (Table 2). No evidence of community
approaches was observed (RR: 0.95; 95% CI: 0.88–1.04) when pooling four studies ($P = 0.121$ for heterogeneity) with a density of > 9 facilities per 100,000 population. However, we found a 16% reduction (RR: 0.84; 95% 0.78–0.90) when pooling five studies ($P < 0.001$ for heterogeneity) in areas with ≤8 facilities per 100,000 population. Only six trials reported on the number of nurses and midwives in the area and we observed no difference in the effect on neonatal mortality in settings with higher or lower number of nurses and midwives per population (Table 2).

**Implementation characteristics**

The mean improvement in immediate breastfeeding was a +29% point change in very high mortality settings, while a change of only +8% points was observed in moderately high mortality settings (Fig. 7). The change in facility births was +6% points (range: −6 to 15) in very high mortality settings, +10% points (range: 0 to 18) in high mortality settings and +3% points (range: −1 to 9) in moderately high mortality settings (Fig. 8).

We observed no evidence that the effect of the community-based approaches on neonatal mortality was associated with improvements in immediate breastfeeding and facility births. The analysis pooling five trials which achieved improvement in immediate breastfeeding of 25% or more suggested a reduction of neonatal mortality of 18% (RR: 0.82; 95% CI: 0.76–0.89; $P < 0.001$ for heterogeneity). Similarly, the pooled analysis of four trials achieving only marginal improvement (≤5%) in immediate breastfeeding suggested a 19% reduction in neonatal mortality (RR: 0.81; 95% CI: 0.74–0.89; $P < 0.001$ for heterogeneity; Table 2).

All home-based counselling interventions reached more than 40% of pregnant women and the size of the effect of the intervention on neonatal mortality did not differ in relation to the proportion of women reached. However, a difference was seen when running a sub-analysis of the women’s group interventions. Pooling four trials that reached 37–66% of pregnant women we found a 26% reduction in neonatal mortality (RR: 0.74; 95% CI: 0.68–0.82; $P = 0.418$ for heterogeneity). In contrast, pooling the five studies which received <36% coverage suggested a lower effect size on neonatal mortality (RR: 0.87; 95% CI: 0.81–0.95; $P < 0.001$ for heterogeneity; Fig. 9; Table 2).
Discussion
Our analysis suggests that large gains in neonatal survival can be achieved using community approaches in settings with very high neonatal mortality and very low rates of facility births. Where mortality is lower, although still moderately high, no evidence of an effect of community approaches on neonatal mortality was found. The observed effect size of the community approaches was larger in south Asia, while there was no evidence of an effect when pooling the studies done in sub-Saharan Africa. This might be partly explained by the fact that the trials in Ghana, Malawi and the United Republic of Tanzania were done in settings with moderately high neonatal mortality.

The large effect of a 45% reduction of neonatal mortality which was previously reported\textsuperscript{5} could be because these early trials were done in settings with high mortality and unhealthy home-care practices. Except in one trial,\textsuperscript{25} subsequent meta-analysis\textsuperscript{9,10} included trials done in places where neonatal mortality was considerably lower.

As neonatal mortality in an area decreases, the relative importance of infectious diseases and other more easily addressable risk factors, such as cold injuries, reduces. The latest work of the Global Burden of Disease group clearly highlights the increasing importance of intrapartum complications, including neonatal encephalopathy, as causes of death.\textsuperscript{31} As non-infectious causes of neonatal mortality become more prominent, health system constraints to prevent intrapartum-related complications and mitigate the effect of prematurity might become more important.\textsuperscript{32} One study concluded that part of the reason their intervention did not result in mortality reduction – despite improved neonatal care and facility coverage – was the failure to address birth asphyxia and prematurity.\textsuperscript{15} A similar argument was raised by others reporting on community and participatory women’s group approaches.\textsuperscript{21,33}

Lower neonatal mortality is likely to reflect recent or ongoing trends in health service uptake, household wealth education and health literacy. One study reported a decrease in neonatal mortality in both intervention and comparison groups, accompanied by increases in newborn care practices and health-service uptake, suggesting underlying trends that had a larger impact than the trial intervention itself.\textsuperscript{20} Others reported a doubling of facility deliveries during the trial period that was possibly due in part to increased transportation and better communications in the area.\textsuperscript{9} A third study suggested that a reduction in neonatal mortality in both intervention and control groups was likely related to improvements in the living environment in the slum areas, such as covering gutters and better sanitation and electricity supplies.\textsuperscript{21} These welcome investments in the health system and overall development reflect a
rapidly changing context in which it is inherently more difficult to show large mortality reductions from specific interventions.\textsuperscript{34}

Many of the trials in our analysis reported only a very modest improvement in the numbers of women delivering in a health-care facility. This is in contrast to the most recent large increases in facility births observed in many low- and middle-income countries; these have occurred because of multiple factors, both within and outside their health systems.\textsuperscript{35}

Going forward, the strategies and content of community approaches to neonatal survival might need to be re-examined. Still many potential benefits of community approaches to enhancing health literacy, reducing delays in care-seeking and improving linkages between the community and health facilities for emergency referral exist. Community approaches can also encourage accountability measures that could support facility strengthening.\textsuperscript{36} However, the effect on neonatal mortality would depend on the quality of services available, and the two effects could not be separated.

Our approach of examining effects of community approaches in relation to context and health-system factors has to be interpreted with caution. We hypothesize that in settings with lower neonatal mortality, more facility births and improved newborn care practices, these interventions may have less effect. However, our stratified meta-analysis cannot prove such an association. Our findings are plausible against the background that the present community approaches target neonatal sepsis and complications of prematurity, while in a context of reduced neonatal mortality, intrapartum-related complications leading to asphyxia become more important.\textsuperscript{37} Reducing intrapartum complications and birth asphyxia will demand quality intrapartum services which the present community approaches do not address.

We combined the two different approaches of home-based counselling and women’s groups in our analysis, which strictly speaking prohibits any meta-analysis. Nevertheless, both approaches aimed to improve mothers’ newborn care practices at home and health-seeking behaviours, and thus the mediators through which they affect neonatal mortality are expected to be similar. Our main aim was not to present summary estimates of the mortality effect to guide policy changes. Rather, we hope to contribute to the development of a theory underpinning the opportunities and limitations of community approaches and the role these approaches might play in the development of care packages to address neonatal mortality in the SDG era.
We computed summary baseline rates of immediate breastfeeding and facility births as well as percentage point changes due to the interventions. However, some authors did not include such information in their papers. While some reported baseline data, others only reported comparisons at endline. As a result the difference-in-difference between intervention and comparison groups could not be calculated. Thus improvements in newborn care practices are not adjusted for differences in baseline values between intervention and comparison groups for some studies. Moreover, our analysis was constrained by the lack of reporting on health system factors such as availability of health facilities or health providers. This reminds us of the value of a careful description of the context in which interventions are implemented to enable an understanding of the transferability of results.

In conclusion, the findings suggest that beyond a certain mortality threshold, community approaches alone might not lead to marked improvements in survival. This finding supports the recent trend in the SDG era towards increasing investment in the quality of facility care.
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### Table 1: Populations, intervention characteristics and intermediate outcomes for randomized cluster studies included in the meta-analysis of community-based approaches for neonatal survival

| Study type and authors | Evaluation period | Area, country | Setting | Neonatal deaths in trial area, per 1000 live births | Study population, no. | Study design | Intervention | No. of clusters in trial | Immediate breastfeeding, % of births | Facility births, % of births | Intermediate outcomes |
|-----------------------|-------------------|---------------|---------|--------------------------------------------------|-----------------------|--------------|--------------|--------------------------|-----------------------------------|-----------------------------|---------------------|
|                        |                   |               |         |                                                  |                       |              |              |                          |                                    |                             | Baseline | Change | Baseline | Change | |
| Home-based counselling trials |                   |               |         |                                                  |                       |              |              |                          |                                    |                             |         |        |         |        | |
| Baqui et al., 200813  | 2003–2005         | Sylhet, Bangladesh | Poor rural | 44 | ~480 000 | Community meetings + home-based counselling visits (2 in pregnancy and 3 postpartum) + home treatment if referral failed | 16 | 45 | +28 | 10 | +1 | |
| Kumar et al., 200814  | 2003–2005         | Shivgarh, India | Poor rural | 84 | 104 123 | Community meetings + home-based counselling visits (2 in pregnancy and 2 postpartum) | 26 | 4 | +65 | 8 | +9 | |
| Kumar et al., 200814  | 2003–2005         | Shivgarh, India | Poor rural | 84 | 104 123 | Community meetings + home-based counselling visits (2 in pregnancy and 2 postpartum) + ThermoSpot | 26 | 3 | +63 | 3 | +15 | |
| Darmstadt et al., 201015 | 2005–2006         | Mirzapur, Bangladesh | Poor rural | 28 | 292 000 | Home-based counselling visits (2 in pregnancy and 4 postpartum) | 12 | 41 | +25 | 12 | +4 | |
| Bhutta et al., 201116  | 2006–2008         | Hala, Pakistan | Poor rural | 49 | 600 000 | Community mobilization + home-based counselling visits (2 in pregnancy and 2 postpartum) | 16 | 27 | +16 | 44 | +10 | |
| Bhandari et al., 201217 | 2008–2010         | Haryana, India | Poor rural | 43 | 1 100 000 | Home-based counselling visits (3 postpartum) | 18 | 11d | +30 | N/A | N/A | |
| Kirkwood et al., 201310 | 2008–2009         | Newhints, Ghana | Poor rural | 32 | 600 000 | Home-based counselling visits (2 in pregnancy and 3 postpartum) | 98 | 41d | +7 | 58 | 0 | |
| Hanson et al., 20159   | 2010–2013         | Mtwara and Lindi, United Republic of Tanzania | Poor rural | 30 | 1 200 000 | Home-based counselling visits (3 in pregnancy and 2 postpartum) | 132 | 19 | +7 | 43 | +2 | |
| Women's group trials | Start-End Year | Location | Setting | Population | Intervention | Duration | N/A | Immediate Breastfeeding | Facilitated Births |
|----------------------|----------------|----------|---------|------------|--------------|----------|-----|------------------------|-------------------|
| Manandhar et al., 2004 | 2001–2003 | Makwanpur, Nepal | Poor rural | 37 | 400 000 | Monthly participatory women’s group meetings | 24 | 54 | +8 | 2 | +5 |
| Tripathy et al., 2010 | 2005–2008 | Jharkhand and Orissa, India | Poor rural | 60 | 228 186 | Monthly participatory learning + action cycle | 36 | 61 | +3 | 20 | −6 |
| Azad et al., 2010 | 2005–2007 | Bogra, Bangladesh | Poor rural | 38 | 503 163 | Participatory learning + action cycle | 18 | 51 | N/A | 7 | 0 |
| More et al., 2012 | 2006–2009 | Mumbai, India | Urban slum | 11 | 282 000 | Bi-monthly participatory meetings including peer learning | 48 | 82 | 0 | 87 | −1 |
| Colbourn et al., 2013 | 2007–2010 | MaiKanda, Malawi | Poor rural | 34 | 2 500 000 | Monthly participatory learning + action cycle | 32 | N/A | N/A | 41 | +17 |
| Colbourn et al., 2013 | 2007–2010 | Kasungu, Lilongwe and Salima, Malawi | Poor rural | 34 | 2 500 000 | Participatory learning + action cycle + facility strengthening | 30 | N/A | N/A | 52 | +18 |
| Fottrell et al., 2013 | 2009–2011 | Bogra, Bangladesh | Poor rural | 30 | 532 996 | Monthly participatory learning + action cycle | 18 | 65 | +7 | 19 | +1 |
| Lewycka et al., 2013 | 2004–2010 | MaiMwana, Malawi | Poor rural | 30 | 185 888 | Monthly participatory learning + action cycle, with and without volunteer peer counselling | 36 | 78 | +2 | 36 | +9 |
| Tripathy et al., 2016 | 2009–2012 | Jharkhand and Orissa, India | Poor rural | 63 | 156 519 | Monthly participatory learning + action cycle | 30 | 77 | +1 | 48 | +4 |

N/A: not available.

a All studies were cluster randomized trials comparing neonatal mortality in the population receiving the intervention with mortality in a comparison population receiving the local standard care.

b Immediate breastfeeding was defined in most studies as the percentage of births in which the infant was breastfed within 1 hour of delivery (mother’s report), except Bhutta et al.16 who defined breastfeeding within 30 minutes, and Tripathy et al.19 who defined breastfeeding within 4 hours of birth. Facility birth was defined in all studies as the percentage of births in a health-care facility. Baseline was the value at the trial baseline in the intervention and comparison groups. Change was the change in values between the trial baseline and endline separately for intervention and comparison groups (the difference-in-differences).

c ThermoSpot™ (Camborne Consultants, Dorset, England) is a non-invasive liquid crystal indicator for hypothermia.

d For trials that did not report on newborn practices at baseline and endline we did not calculate the difference-in-difference change but the simple difference between estimates from intervention and comparison group.

e The intervention group included all 24 clusters with women’s groups (with and without additional peer counselling). The comparison group included the 12 clusters without any intervention. However, the data on increases in breastfeeding and facility births were calculated with the comparison group of all clusters without women’s groups as no data were available separately for the clusters with no intervention.

Note: This table shows intermediate outcomes but the primary outcome for all studies was population-based neonatal mortality rate obtained either from surveys or continuous surveillance in the target population.
Table 2. Effect on neonatal mortality of trials of community-based approaches for neonatal survival, stratified by context and implementation characteristics

| Stratification variable | No. of trials or trial arms | RR (95% CI) random effects model | Tests for heterogeneity/$I^2$, % | P for heterogeneity in sub-groups |
|-------------------------|----------------------------|--------------------------------|-------------------------------|----------------------------------|
| Neonatal mortality in comparison group, no. of deaths per 1000 live births | | | | |
| ≤ 32 | 6 | 0.94 (0.88–1.01) | 82 | < 0.001 |
| 33–43 | 5 | 0.89 (0.83–0.95) | 3 | 0.392 |
| ≥ 44 | 6 | 0.75 (0.69–0.80) | 73 | 0.002 |
| Geographical area | | | | |
| South Asia | 12 | 0.82 (0.78–0.86) | 81 | < 0.001 |
| Sub-Saharan Africa | 5 | 0.95 (0.88–1.02) | 34 | 0.193 |
| Immediate breastfeeding at baseline, % of births$^{a,b}$ | | | | |
| ≤ 25 | 5 | 0.91 (0.85–0.98) | 87 | < 0.001 |
| 26–53 | 4 | 0.87 (0.81–0.94) | 29 | 0.239 |
| ≥ 54 | 5 | 0.81 (0.73–0.90) | 85 | < 0.001 |
| Facility births at baseline, % of births$^{b,c}$ | | | | |
| ≤ 10 | 5 | 0.77 (0.71–0.85) | 80 | 0.001 |
| 11–43 | 6 | 0.85 (0.80–0.91) | 80 | < 0.001 |
| ≥ 44 | 5 | 0.90 (0.83–0.97) | 80 | 0.001 |
| Density of facilities in study area, no. per 100 000 population | | | | |
| ≤ 8 | 5 | 0.84 (0.78–0.90) | 74 | < 0.001 |
| > 9 | 4 | 0.95 (0.88–1.04) | 48 | 0.121 |
| Density of nurses and midwives in study area, no. per 1000 population | | | | |
| ≤ 0.4 | 4 | 0.85 (0.79–0.92) | 87 | < 0.001 |
| > 0.4 | 2 | 0.86 (0.73–0.99) | 0 | 0.721 |
| Type of intervention | | | | |
| Home-based counselling | 8 | 0.89 (0.85–0.94) | 80 | < 0.001 |
| Women’s group | 9 | 0.82 (0.77–0.87) | 75 | < 0.001 |
| Immediate breastfeeding, % points change at endline$^d$ | | | | |
| ≤ +5 | 4 | 0.81 (0.74–0.89) | 88 | < 0.001 |
| +5 to +24 | 5 | 0.90 (0.84–0.96) | 79 | 0.001 |
| ≥ +25 | 5 | 0.82 (0.76–0.89) | 83 | < 0.001 |
| Facility births, % points change at endline$^d$ | | | | |
| ≤ +1 | 6 | 0.83 (0.78–0.88) | 84 | < 0.001 |
| +2 to +8 | 4 | 0.92 (0.85–1.00) | 79 | < 0.003 |
| ≥ +9 | 6 | 0.81 (0.75–0.88) | 73 | 0.002 |
| Coverage of home-based counselling, % of pregnant women |   |     |     |     |
|--------------------------------------------------------|---|-----|-----|-----|
| 37–66                                                  | 3 | 0.92 (0.86–0.99) | 81  | 0.005 |
| ≥ 67                                                   | 5 | 0.86 (0.79–0.93) | 83  | < 0.001 |

| Coverage of women's groups, % of pregnant women attending | |     |     |     |
|------------------------------------------------------------|---|-----|-----|-----|
| ≤ 36                                                       | 5 | 0.87 (0.81–0.95) | 83  | < 0.001 |
| 37–66                                                      | 4 | 0.74 (0.68–0.82) | 0   | 0.418 |

CI: confidence interval; RR: relative risk.

| Note | Description |
|------|-------------|
| a    | Immediate breastfeeding was defined in most studies as the percentage of births in which the infant was breastfed within 1 hour of delivery, except Bhutta et al.\(^1\) who defined breastfeeding within 30 minutes, and Tripathy et al.\(^2\) who defined breastfeeding within 4 hours of birth. |
| b    | Baseline was the value at the trial baseline (in the intervention and comparison groups). |
| c    | Facility birth was defined in all studies as the percentage of births in a health-care facility. |
| d    | Change was the change in values between the trial baseline and endline separately for intervention and comparison groups (the difference-in-differences). |
| e    | Percentage of pregnant women visited at home by a community health worker. |
| f    | Percentage of pregnant women attending their local women's group. |
Fig. 1. Flowchart showing the selection of articles for meta-analysis of the effect of community approaches for neonatal survival

**Home-based counselling trials**

- Search by Gogia & Sachdev 2010 (PubMed, the Cochrane Controlled Trials Register, EMBASE, HealthSTAR and CINAHL); no search terms provided
- 5 trials (2 randomized):
  - Bang et al., 2005
  - Baqui et al., 2008\(^a\)
  - Baqui et al., 2008\(^b\)
  - Bhutta et al., 2008
  - Kumar et al., 2008
- Search by Kirkwood et al. 2013
  - Databases searched not provided
  - Search terms: “newborn”, “neonatal”, “mortality”, “cluster” & “trial”
  - Exclusion of 1 trial: Baqui et al., 2008\(^b\)
- 4 proof-of-principle trials (as above), plus 4 trials:
  - Darmstadt et al., 2010
  - Bhutta et al., 2011
  - Bhandari et al., 2012
  - Kirkwood et al., 2013
- Search by Hansen et al., 2015; Search terms: (((“newborn” OR “neonatal” OR “maternal”)) AND mortality) AND trial
  - Exclusion of 2 non-randomized trials:
    - Bang et al., 2005
    - Bhutta et al., 2008

**Women’s group trials**

- Search by Prost et al., 2013
  - Search terms: “community mobilization”, “community participation”, “participatory action”, “participatory learning and action”, “women’s group” & “women”
- 7 trials:
  - Manadhar et al., 2004
  - Azad et al., 2010
  - Tripathy et al., 2010
  - Moe et al., 2012
  - Celbourn et al., 2013
  - Fottrell et al., 2013
  - Lewycke et al., 2013
- Searched PubMed and Web of Science online databases using search string (((“newborn” OR “neonatal” OR “maternal”)) AND mortality) AND trial
  - Identified and screened 1481 titles published between Jan 2013 and May 2016
- 7 papers included in the analysis of meta-analysis of Prost et al., 2013 plus Tripathy et al., 2016

\(^a\)Baqui et al., 2008

\(^b\)Baqui et al., 2008
Fig. 2. Meta-analysis of the effect on neonatal mortality of trials of community approaches for neonatal survival, by neonatal mortality rate at baseline

| Author          | Country       | Evaluation period | Neonatal deaths in trial area, per 1000 live births | RR (95% CI)     | % weight |
|-----------------|---------------|-------------------|-----------------------------------------------|-----------------|----------|
| ≥ 44 deaths per 1000 live births |                |                   |                                               |                 |          |
| Kumar 2008A     | India         | 2003–2005         | 84                                            | 0.50 (0.37–0.68) | 2.2      |
| Kumar 2008B     | India         | 2003–2005         | 84                                            | 0.53 (0.37–0.74) | 1.9      |
| Sripathy 2016   | India         | 2009–2012         | 63                                            | 0.68 (0.54–0.87) | 3.1      |
| Sripathy 2010   | India         | 2005–2008         | 60                                            | 0.73 (0.64–0.82) | 11.0     |
| Bhutta 2011     | Pakistan      | 2006–2008         | 49                                            | 0.88 (0.78–0.99) | 11.4     |
| Bago 2008       | Bangladesh    | 2003–2005         | 44                                            | 0.72 (0.59–0.88) | 4.3      |
| Subtotal (I-squared = 73%, P = 0.002) | |                   |                                               | 0.75 (0.69–0.80) | 33.9     |

| 33–43 deaths per 1000 live births |                |                   |                                               |                 |          |
| Bhandari 2012   | India         | 2008–2010         | 43                                            | 0.91 (0.84–1.00) | 14.0     |
| Zulu 2010       | Malawi        | 2005–2007         | 38                                            | 0.90 (0.80–1.01) | 11.4     |
| Maranathini 2004 | Nepal        | 2001–2003         | 37                                            | 0.71 (0.54–0.94) | 2.3      |
| Colbourn 2003B  | Malawi        | 2007–2010         | 34                                            | 0.79 (0.64–0.99) | 3.4      |
| Colbourn 2003A  | Malawi        | 2007–2010         | 34                                            | 0.88 (0.71–1.09) | 3.3      |
| Subtotal (I-squared = 3%, P = 0.392) |                 |                   |                                               | 0.89 (0.83–0.95) | 34.5     |

| ≤ 32 deaths per 1000 live births |                |                   |                                               |                 |          |
| Kirkwood 2013   | Ghana         | 2008–2009         | 32                                            | 0.99 (0.78–1.21) | 5.0      |
| Harnon 2015     | United Republic of Tanzania | 2010–2013 | 30                                            | 1.03 (0.93–1.15) | 12.7     |
| Dewycza 2013    | Malawi        | 2004–2010         | 30                                            | 0.86 (0.71–1.05) | 3.9      |
| Fettell 2013    | Bangladesh    | 2009–2011         | 30                                            | 0.67 (0.55–0.81) | 5.5      |
| Drumstee 2010   | Bangladesh    | 2005–2006         | 28                                            | 0.86 (0.68–1.01) | 2.8      |
| Mohe 2012       | India         | 2006–2009         | 11                                            | 1.47 (1.12–1.92) | 1.8      |
| Subtotal (I-squared = 82%, P < 0.001) |                 |                   |                                               | 0.94 (0.88–1.01) | 31.7     |
| Overall (I-squared = 79%, P < 0.001) |                 |                   |                                               | 0.86 (0.82–0.89) | 100.0    |

CI: confidence interval; RR: relative risk.

Notes: For references with more than one trial, each trial is presented separately and denoted with a letter after the date. Baseline neonatal mortality rate was defined as the number of deaths within 28 days after birth per 1000 live births in the comparison group in the study area.
Fig. 3. Meta-analysis of the effect on neonatal mortality of trials of community approaches for neonatal survival, by type of approach

| Author               | Country          | Evaluation period | Neontal deaths in trial area, per 1000 live births | RR (95% CI)     | % weight |
|----------------------|------------------|-------------------|--------------------------------------------------|-----------------|----------|
| **Home-based counselling** |                  |                   |                                                  |                 |          |
| Kumar 2008A          | India            | 2003–2005         | 84                                               | 0.50 (0.37–0.68) | 2.2      |
| Kumar 2008B          | India            | 2003–2005         | 84                                               | 0.52 (0.37–0.74) | 1.9      |
| Bhutta 2011          | Pakistan         | 2006–2008         | 49                                               | 0.88 (0.78–0.99) | 11.4     |
| Baqui 2006           | Bangladesh       | 2003–2005         | 44                                               | 0.72 (0.59–0.88) | 4.3      |
| Bhattacharjee 2012   | India            | 2008–2010         | 43                                               | 0.93 (0.84–1.03) | 14.1     |
| Kirkwood 2013        | Ghana            | 2008–2009         | 32                                               | 0.93 (0.78–1.11) | 5.0      |
| Hanson 2015          | United Republic of Tanzania | 2010–2013  | 30                                               | 1.03 (0.93–1.15) | 12.7     |
| Dammstadt 2010       | Bangladesh       | 2005–2006         | 28                                               | 0.86 (0.68–1.10) | 2.8      |
| **Subtotal (I-squared = 80%, P < 0.001)** |                  |                   |                                                  | 0.89 (0.85–0.94) | 54.3     |
| **Women's group**    |                  |                   |                                                  |                 |          |
| Tripathy 2016        | India            | 2009–2012         | 63                                               | 0.68 (0.54–0.87) | 3.1      |
| Tripathy 2010        | India            | 2005–2008         | 60                                               | 0.71 (0.64–0.82) | 11.0     |
| Acad 2010            | Bangladesh       | 2005–2007         | 38                                               | 0.90 (0.80–1.01) | 11.4     |
| Manandhar 2004       | Nepal            | 2001–2003         | 37                                               | 0.71 (0.54–0.94) | 2.3      |
| Colbourne 2013B      | Malawi           | 2007–2010         | 34                                               | 0.79 (0.64–0.99) | 3.4      |
| Colbourne 2013A      | Malawi           | 2007–2010         | 34                                               | 0.88 (0.71–1.10) | 3.3      |
| Lwewycz 2013         | Malawi           | 2004–2010         | 30                                               | 0.86 (0.71–1.06) | 3.9      |
| Forrester 2013       | Bangladesh       | 2009–2011         | 30                                               | 0.67 (0.56–0.81) | 5.5      |
| More 2012            | India            | 2006–2009         | 11                                               | 1.42 (1.12–1.82) | 1.8      |
| **Subtotal (I-squared = 75%, P < 0.001)** |                  |                   |                                                  | 0.73 (0.57–0.97) | 45.7     |
| **Overall (I-squared = 78%, P < 0.001)** |                  |                   |                                                  | 0.86 (0.82–0.89) | 100.0    |

CI: confidence interval; RR: relative risk.

*a Baqui et al.13*

Notes: For references with more than one trial, each trial is presented separately and denoted with a letter after the date. Home-based behaviour-change counselling involved home visits to individual pregnant women by a community health worker and sometimes included community meetings. Women’s participatory groups took place in the community and were facilitated by trained community members. Both approaches included education, behaviour change communication and a problem-solving approach to improve newborn care behaviours by mothers, such as immediate and exclusive breastfeeding, thermal care, and safe and dry cord care.
Fig. 4. Meta-analysis of the effect on neonatal mortality of trials of community approaches for neonatal survival, by region

| Author        | Country     | Evaluation period | RR (95% CI) | % weight |
|---------------|-------------|-------------------|-------------|----------|
| South Asia    |             |                   |             |          |
| Kumar 2008A  | India       | 2003–2005         | 0.50 (0.37–0.68) | 2.2      |
| Kumar 2008B  | India       | 2003–2005         | 0.52 (0.37–0.74) | 1.9      |
| Tripathy 2016 | India       | 2009–2012         | 0.68 (0.54–0.87) | 3.1      |
| Tripathy 2010 | India       | 2005–2008         | 0.73 (0.64–0.82) | 11.0     |
| Bhutta 2011   | Pakistan    | 2006–2008         | 0.88 (0.78–0.99) | 11.4     |
| Baqui 2008†   | Bangladesh  | 2003–2005         | 0.72 (0.59–0.88) | 4.3      |
| Bhandari 2012 | India       | 2006–2010         | 0.95 (0.84–1.03) | 14.0     |
| Areal 2010    | Bangladesh  | 2005–2007         | 0.91 (0.80–1.01) | 11.4     |
| Manandhar 2004 | Nepal     | 2001–2003         | 0.71 (0.54–0.94) | 2.3      |
| Fettweis 2013 | Bangladesh  | 2009–2011         | 0.67 (0.56–0.81) | 5.5      |
| Dumont 2010   | Bangladesh  | 2005–2006         | 0.86 (0.68–1.10) | 2.8      |
| More 2012     | India       | 2006–2009         | 1.47 (1.12–1.95) | 1.8      |
| Subtotal (I-squared = 81%, P < 0.001) |            |                   | 0.82 (0.79–0.86) | 71.1      |
| Sub-Saharan Africa |          |                   |             |          |
| Celloura 2013A | Malawi     | 2007–2010         | 0.79 (0.64–0.99) | 3.4      |
| Celloura 2013A | Malawi     | 2007–2010         | 0.88 (0.71–1.10) | 3.3      |
| Kirkwood 2013 | Ghana      | 2006–2009         | 0.93 (0.79–1.11) | 5.0      |
| Hansen 2015   | United Republic of Tanzania | 2010–2013 | 1.05 (0.91–1.15) | 12.7 |
| Lewin 2013    | Malawi     | 2004–2010         | 0.86 (0.74–1.00) | 3.9      |
| Subtotal (I-squared = 34%, P = 0.193) |            |                   | 0.95 (0.88–1.02) | 28.3     |
| Overall (I-squared = 79%, P < 0.001) |            |                   | 0.86 (0.82–0.89) | 100.0    |

CI: confidence interval; RR: relative risk.

† Baqui et al.13

Note: For references with more than one trial, each trial is presented separately and denoted with a letter after the date.
Fig. 5. Meta-analysis of the effect on neonatal mortality of trials of community approaches for neonatal survival, by immediate breastfeeding at baseline.

| Author       | Country          | Immediate breastfeeding, % coverage | RR (95% CI) | % weight |
|--------------|------------------|-------------------------------------|-------------|----------|
| ≤ 25% immediate breastfeeding |
| Kumar 2008B  | India            | 3                                   | 0.52 (0.37–0.74) | 2.2      |
| Kumar 2008A  | India            | 4                                   | 0.50 (0.37–0.68) | 2.6      |
| Bhundari 2012| India            | 11                                  | 0.93 (0.84–1.03) | 17.1     |
| Hanson 2015  | United Republic of Tanzania | 19                                | 1.03 (0.93–1.15) | 15.4     |
| Darmstad 2010| Bangladesh       | 25                                  | 0.86 (0.68–1.10) | 3.4      |
| Subtotal (I-squared = 47%, P < 0.001) |                |                                     | 0.91 (0.85–0.98) | 40.7     |
| 26–53% immediate breastfeeding |
| Bhutta 2011  | Pakistan         | 27                                  | 0.88 (0.78–0.99) | 13.9     |
| Kirkwood 2013| Ghana           | 41                                  | 0.93 (0.78–1.11) | 6.1      |
| Baqui 2008A  | Bangladesh       | 45                                  | 0.72 (0.59–0.88) | 5.3      |
| Azad 2010    | Bangladesh       | 51                                  | 0.90 (0.80–1.01) | 13.9     |
| Subtotal (I-squared = 39%, P = 0.239) |                |                                     | 0.87 (0.81–0.94) | 39.1     |
| ≥ 54% immediate breastfeeding |
| Alexandhar 2004 | Nepal          | 54                                  | 0.71 (0.54–0.94) | 2.8      |
| Fottrell 2013 | Bangladesh       | 65                                  | 0.67 (0.56–0.81) | 6.7      |
| Tripathy 2016 | India           | 77                                  | 0.68 (0.54–0.87) | 3.8      |
| Lewycka 2013 | Malawi           | 78                                  | 0.66 (0.51–1.05) | 4.2      |
| More 2012    | India            | 82                                  | 1.47 (1.12–1.92) | 2.2      |
| Subtotal (I-squared = 85%, P < 0.001) |                |                                     | 0.81 (0.73–0.90) | 20.2     |
| Overall (I-squared = 89%, P < 0.001) |                |                                     | 0.86 (0.82–0.90) | 100.0    |

CI: confidence interval; RR: relative risk.

* Baqui et al.13

Notes: For references with more than one trial, each trial is presented separately and denoted with a letter after the date. Immediate breastfeeding was defined in most studies as the percentage of births in which the infant was breastfed within 1 hour of delivery, except Bhutta et al.16 who defined breastfeeding within 30 minutes, and Tripathy et al.19 who defined breastfeeding within 4 hours of birth.
Fig. 6. Meta-analysis of the effect on neonatal mortality of trials of community approaches for neonatal survival, by facility births at baseline

| Author            | Country       | % coverage | RR (95% CI) | % weight |
|-------------------|---------------|------------|-------------|----------|
| ≤ 10% facility birth |               |            |             |          |
| Mehanchar 2004    | Nepal         | 2          | 0.71 (0.54–0.94) | 2.6      |
| Kamar 2008B       | India         | 3          | 0.92 (0.37–0.74) | 2.2      |
| Azad 2010         | Bangladesh    | 7          | 0.90 (0.20–1.01) | 13.3     |
| Kamar 2008A       | India         | 8          | 0.50 (0.37–0.68) | 2.5      |
| Baqui 2008A       | Bangladesh    | 10         | 0.72 (0.59–0.88) | 5.0      |
| Subtotal (I² = 80%, P < 0.001) |     |            | 0.77 (0.71–0.85) | 25.6     |
| 11%–43% facility birth |            |            |             |          |
| Damstrok 2010     | Bangladesh    | 12         | 0.86 (0.68–1.10) | 3.2      |
| Fottrell 2013     | Bangladesh    | 19         | 0.67 (0.56–0.81) | 6.4      |
| Tripathy 2010     | India         | 20         | 0.73 (0.64–0.82) | 12.8     |
| Lewycke 2013      | Malawi        | 36         | 0.86 (0.71–1.06) | 4.5      |
| Golburn 2013A     | Malawi        | 41         | 0.88 (0.71–1.10) | 3.9      |
| Hanson 2015       | United Republic of Tanzania | 43 | 1.03 (0.93–1.15) | 14.8     |
| Subtotal (I² = 80%, P < 0.001) |     |            | 0.85 (0.80–0.91) | 45.6     |
| ≥ 44% facility birth |            |            |             |          |
| Bhutta 2011       | Pakistan      | 44         | 0.88 (0.78–0.99) | 13.3     |
| Tripathy 2016     | India         | 48         | 0.68 (0.54–0.87) | 3.6      |
| Colburn 2013B     | Malawi        | 52         | 0.79 (0.64–0.95) | 3.9      |
| Kikekwo 2013      | Ghana         | 58         | 0.93 (0.78–1.11) | 5.9      |
| Mone 2012         | India         | 87         | 1.47 (1.32–1.92) | 2.1      |
| Subtotal (I² = 80%, P = 0.001) |     |            | 0.90 (0.83–0.97) | 28.8     |
| Overall (I² = 79%, P < 0.001) |        |            | 0.85 (0.81–0.88) | 100.0    |

CI: confidence interval; RR: relative risk.

Notes: For references with more than one trial, each trial is presented separately and denoted with a letter after the date. Facility birth was defined in all studies as the percentage of births in a health-care facility.
Fig. 7. Mean baseline and changes in proportion of women breastfeeding immediately after delivery, by neonatal mortality in trial area

Notes: Immediate breastfeeding was defined in most studies as the percentage of births in which the infant was breastfed within 1 hour of delivery, except Bhutta et al.\textsuperscript{16} who defined breastfeeding within 30 minutes, and Tripathy et al.\textsuperscript{19} who defined breastfeeding within 4 hours of birth.

$n$ is the number of trials.

Fig. 8. Mean baseline and changes in proportion of women delivering in a facility, by neonatal mortality in trial area

Note: Facility birth was defined in all studies as the percentage of births in a health-care facility.

$n$ is the number of trials.
Fig. 9. Meta-analysis of the effect on neonatal mortality of community approaches for neonatal survival in women’s group trials, by coverage of pregnant women

| Author       | Country  | Study type | Coverage of pregnant women | RR (95% CI) | % weight |
|--------------|----------|------------|----------------------------|-------------|----------|
| ≤ 36% coverage of pregnant women |
| Azzad 2016   | Bangladesh | Women's group | 3  | 0.99 (0.88–1.01) | 25.0 |
| Kolbourn 2013B | Malawi   | Women's group | 10 | 0.79 (0.64–0.99) | 7.4  |
| Kolbourn 2013A | Malawi   | Women's group | 10 | 0.88 (0.71–1.10) | 7.3  |
| Fetteli 2013  | Bangladesh | Women's group | 36 | 0.67 (0.56–0.81) | 12.1 |
| Subtotal (I-squared = 83%, P < 0.001) | | | | 0.87 (0.81–0.95) | 55.7 |
| 37–66% coverage of pregnant women |
| Trapha 2010  | India    | Women's group | 37 | 0.73 (0.64–0.82) | 24.1 |
| Manandhar 2004 | Nepal   | Women's group | 37 | 0.71 (0.54–0.94) | 5.0  |
| Levyck 2013  | Malawi   | Women's group | 51 | 0.86 (0.71–1.06) | 8.4  |
| Trapha 2016  | India    | Women's group | 66 | 0.68 (0.54–0.87) | 6.8  |
| Subtotal (I-squared = 0%, P = 0.419) | | | | 0.74 (0.68–0.82) | 44.3 |
| Overall (I-squared = 75%, P < 0.001) | | | | 0.82 (0.77–0.87) | 100.0 |

CI: confidence interval; RR: relative risk.

Notes: For references with more than one trial, each trial is presented separately and denoted with a letter after the date. Coverage of home-based counselling was the percentage of pregnant women visited at home by community health workers. Coverage of women’s groups was the percentage of pregnant women attending their local women’s group.