VERTEX-REINFORCED RANDOM WALK

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ABSTRACT:

This paper considers a class of non-Markovian discrete-time random processes on a finite state space \(\{1, \ldots, d\}\). The transition probabilities at each time are influenced by the number of times each state has been visited and by a fixed a priori likelihood matrix, \(R\), which is real, symmetric and nonnegative. Let \(S_i(n)\) keep track of the number of visits to state \(i\) up to time \(n\), and form the fractional occupation vector, \(V(n)\), where \(v_i(n) = S_i(n)/(\sum_{j=1}^{d} S_j(n))\). It is shown that \(V(n)\) converges to to a set of critical points for the quadratic form \(H\) with matrix \(R\), and that under nondegeneracy conditions on \(R\), there is a finite set of points such that with probability one, \(V(n) \to p\) for some \(p\) in the set. There may be more than one \(p\) in this set for which \(P(V(n) \to p) > 0\). On the other hand \(P(V(n) \to p) = 0\) whenever \(p\) fails in a strong enough sense to be maximum for \(H\).

Key words: random walk, reinforcement, unstable equilibria, strong law

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1 Introduction

This paper considers a stochastic process in discrete time on a finite state space \( \{1, \ldots, d\} \), in which the probability of a transition to site \( j \) increases each time \( j \) is visited. To define the process, let \( \mathbf{R} \) be a real symmetric \( d \times d \) matrix with \( R_{ij} \geq 0 \) for each \( i, j \), and \( \sum_i R_{ij} > 0 \) for each \( j \). For \( n \geq d \), inductively define random variables \( Y_n \) and \( \mathbf{S}(n) = (S_1(n), \ldots, S_d(n)) \) as follows. Let \( S_i(d) = 1 \) for \( i = 1, \ldots, d \) and let \( Y_d = 1 \). Let \( \mathcal{F}_n \) be the \( \sigma \)-field generated by \( Y_j : d \leq j \leq n \) and let \( Y_{n+1} \) satisfy

\[
P(Y_{n+1} = j \mid \mathcal{F}_n) = R_{Y_n,j} S_j(n) / \sum_i R_{Y_n,i} S_i(n).
\]

Let \( S_i(n+1) = S_i(n) + \delta_{Y_n+1,i} \). In other words, \( \mathbf{S}(n) \) counts one plus the number of times \( Y \) has occupied each state. The sequence of ordered pairs \( (Y_n, \mathbf{S}(n)) \) is a Markov chain, whereas the sequence \( Y_n \) is not.

Define \( \mathbf{V}(n) = \mathbf{S}(n)/n \), so that each \( \mathbf{V}(n) \) is an element of the \( d - 1 \)-simplex \( \Delta \subseteq \mathbb{R}^d \). (In general, boldface is used for vectors and lightface is used for their components.)

This paper studies the question of when \( \mathbf{V}(n) \) converges and to which possible limits. Since \( \mathbf{V}(n) \) may be viewed as an empirical occupation measure for the \( Y \) process, this is essentially asking whether \( Y \) obeys a strong law of large numbers. A few remarks about the model are in order.

The process is meant to model learning behavior. Think of \( R_{ij} \) as a set of initial transition probabilities; each time \( Y \) visits site \( j \), this choice is positively reinforced, resulting in transition probabilities proportional to \( R_{ij} S_j \). The choice of starting state, \( Y_d = 1 \), is arbitrary; also, setting each \( S_i(d) \) equal to one is a matter of convenience and in fact the theorems in this paper are true for any choice of \( S_i(d) > 0 \) and any \( Y_d \in \{1, \ldots, d\} \). The requirement that \( \mathbf{R} \) be symmetric may not always be reasonable in applications, but is essential for our arguments.

Similar models have been studied in [3] under the name of random processes with...
complete connections. When the entries of $R$ are all one, the model reduces to a Pólya urn model; the behavior in this case is atypical, since most of our results apply to the “generic” case where $R$ is invertible. Another similar process called edge-reinforced random walk is studied in [1, 5, 6, 2]; in that case, transitions from $i$ to $j$ are positively reinforced each time a transition is made from $i$ to $j$ or $j$ to $i$. Thinking of the process as traversing a graph with vertices $1, \ldots, d$, this kind of reinforcement keeps track of moves along each edge of the graph, while the process studied in the present paper keeps track of visits to each vertex. Strong laws for edge-reinforced random walk can be found in [1, 5, 2].

The remainder of this introductory section motivates and states the main results. Subsequent sections give proofs of of the four results. Examples and open questions are discussed in the final section.

**Definition 1** For $v \in \triangle$, let $N_i(v) = \sum_j R_{ij}v_i$. Abbreviate this by $N_i$ when a particular vector $v$ may be understood.

**Definition 2** For $v \in \triangle$, let $H(v) = \sum_i v_i N_i(v) = \sum_{ij} R_{ij} v_i v_j$.

**Definition 3** For $v \in \triangle$ such that $H(v) > 0$, define a vector $\pi(v) \in \triangle$ by $\pi_i(v) = v_i N_i(v)/H(v)$.

**Definition 4** For $v \in \triangle$ such that $H(v) > 0$, define a Markov transition matrix $M(v)$ by $M_{ij}(v) = R_{ij} v_j / N_i$.

Note that $H(V(n))$ is below by $\min\{R_{ij} : R_{ij} > 0\}$. Thus $H$ never vanishes on the closure of the set of possible values of $V(n)$, and the clauses about $H$ not vanishing in the above definitions are merely pro forma. For a fixed $v$, $(\pi M)_i = \sum_j \pi_i M_{ij} =$
\[ \sum_j (v_i N_i / H) (R_{ij} v_j / N_i) = \sum_j v_i v_j R_{ij} / N_i = v_i N_i / H = \pi_i, \] so \( \pi(v) \) is an invariant probability for the transition matrix \( M(v) \). The behavior of \( V(n) \) can heuristically be explained as follows.

For \( n \gg L \gg 1 \), compare \( V(n + L) \) to \( V(n) \). Since \( n \gg L \), the \( Y \) process between these times behaves as if \( V \) is not changing, and hence approximates a Markov chain with transition matrix \( M(V(n)) \). Since \( L \gg 1 \), the occupation measure between these times will be close to the invariant measure \( \pi(V(n)) \). This means that \( V(n + L) \approx V(n) + (L/n)(\pi(V(n)) - V(n)) \). Passing to a continuous time limit gives

\[ \frac{d}{dt} V(t) = \frac{1}{t} (\pi(V(t)) - V(t)). \tag{1} \]

Up to an exponential time change, \( V \) should then behave like an integral curve for the vector field \( \pi - I \). One would expect convergence to a critical point or set and, because of the random perturbations, one would not expect convergence to any unstable equilibrium. It is not in general possible to find a potential for this vector field, but the function \( H \) is a Lyapunov function for it. Then one expects convergence of \( V(n) \) to a maximum for \( H \).

**Definition 5** Let \( \mathcal{C} \subseteq \triangle \) be the set of points \( v \) for which \( \pi(v) = v \). The term critical point will be used to denote points of \( \mathcal{C} \). Let \( \mathcal{C}_0 \subseteq \triangle \) be the set of points \( v \) for which \( M(v) \) is reducible.

Section 2 will discuss the nature of \( \mathcal{C} \) and \( \mathcal{C}_0 \), and give conditions under which Theorem 1.1 (proved in Section 3) implies almost sure convergence of \( V(n) \).

**Theorem 1.1** With probability one, \( \text{dist}(V(n), \mathcal{C} \cup \mathcal{C}_0) \to 0 \), where \( \text{dist}(x, A) \) denotes \( \inf\{|x - y| : y \in A\} \).

**Definition 6** For \( v \in \triangle \), define \( \text{face}(v) = \{w \in \triangle : \forall i, v_i = 0 \text{ implies } w_i = 0\} \) to be the closed face of \( \triangle \) to which \( v \) is interior.
**Definition 7** For any \( p \in C \) that is in a proper face of \( \triangle \) a linear non-maximum iff

\[
D_p H(e_k - e_j) > 0 \text{ for some } e_k \notin \text{face}(p), e_j \in \text{face}(p).
\]  

(2)

(Here \( e_1, \ldots, e_d \) are the standard basis vectors in \( \mathbb{R}^d \).)

The following theorems, proved in Section 5 and 4 respectively, give conditions under which convergence to a critical point is impossible.

**Theorem 1.2** Suppose that \( R \) is nonsingular and let \( p \) be the unique critical point in the interior of \( \triangle \). Then \( P(V(n) \to p) = 0 \) whenever \( p \) fails to be a maximum for \( H \). This happens if and only if \( R \) has more than one positive eigenvalue, which happens if and only if the linear operator \( D_p (\pi - I) \) on \( -p + \triangle \) has a positive eigenvalue.

**Theorem 1.3** Suppose \( p \) is a linear non-maximum in a proper face of \( \triangle \). Then \( P(V(n) \to p) = 0. \)

A sort of converse to these nonconvergence theorems gives a criterion for convergence with positive probability of \( V(n) \) to stable critical points. This is proved in Section 3

**Theorem 1.4** Let \( A \) be a component of \( C \) disjoint from \( C_0 \) and suppose that \( A \) is a local maximum for \( H \) in the sense that there is some neighborhood \( N \) of \( A \) for which \( v \in N \) and \( p \in A \) imply \( H(v) < H(p) \). Then \( P(\text{dist}(V(n), A) \to 0) > 0. \)

2 Preliminaries

The following proposition verifies that \( H \) is a Lyapunov function for the vector field \( \pi - I \) and gives alternate characterizations of the set of critical points. The notation used
throughout for vector calculus is $D_v F(w)$ to denote the derivative of $F$ in the direction $w$ at the point $v$, thus $D_v F$ denotes the linear operator approximating $F(v + \cdot) - F(v)$.

**Lemma 2.1** For any $v \in \Delta$, $D_v H(\pi(v) - v) \geq 0$. Furthermore, the following are equivalent:

(i) $D_v H(\pi(v) - v) = 0$

(ii) $D_v H|_{face(v)} = 0$

(iii) for those $i$ such that $v_i > 0$, $N_i$ are equal

(iv) for all $i$, $v_i = \sum_j R_{ij} v_i v_j / N_j$

(v) $\pi(v) = v$

where $0/0 = 0$ in (iv) by convention.

Proof: For fixed $i$ and $j$ and constant $c$, consider the operation of increasing $v_j$ by the quantity $cv_i v_j (N_j - N_i)$ and decreasing $v_i$ by the same amount. When $c = 1/H(v)$ and this operation is done simultaneously for every (unordered) pair $i, j$, then the resulting vector is $\pi(v)$: the next value of the $i^{th}$ coordinate is given by

$$v_i + (1/H(v)) (\sum_j v_i v_j N_i - \sum_j v_i v_j N_j)$$

$$= v_i + (1/H(v)) (v_i N_i - v_i H(v)) = \pi_i(v).$$

So an infinitesimal move towards $\pi(v)$ corresponds to doing these additions and subtractions simultaneously with an infinitesimal $c$. To show that this increases $H$, it suffices to show that for each unordered pair $i, j$, the value of $H$ is increased, since $H$ is smooth and therefore well approximated by its linearization near any point. So let $i, j$ be arbitrary. Writing $v^{(1)}$ for the new vector gives

$$H(v^{(1)}) = \sum_{r,s} R_{rs} v_r^{(1)} v_s^{(1)}$$

$$= \sum_{r,s} R_{rs} v_r v_s + 2 \sum_s R_{is} c v_i v_j (N_i - N_j) v_s$$

$$+ 2 \sum_r R_{rj} c v_i v_j (N_j - N_i)$$

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\[ H(v) + 2cv_iv_j(N_i - N_j)^2 \geq H(v) \]

so \( H \) is nondecreasing. This proves the first part.

For the equivalences, first note that if there are any \( i \) and \( j \) for which \( N_i \neq N_j \) and neither \( v_i \) nor \( v_j \) is zero, then \( H \) strictly increases. Thus \( (i) \iff (iii) \). Since

\[ D_vH \text{ is just inner product with the vector } (2N_1, \ldots, 2N_n), \]  

and restricting to face(\( v \)) just throws out the coordinates \( i \) such that \( v_i = 0 \), it is easy to see that \( (ii) \iff (iii) \). Assuming \( (iii) \), suppose the common value of the \( N_i \) is \( c \). Then multiplying \( (iv) \) by \( c \) gives \( \sum_j v_i v_j = c \cdot v_i \), so \( (iii) \Rightarrow (iv) \). Now assume \( (iv) \). Letting \( M_v \) denote the matrix as well as the Markov chain, \( (iv) \) just says that \( v \) is stationary for \( M_v \). Then \( \pi(v) - v = \vec{0} \) so \( (v) \) holds. And finally, \( (v) \Rightarrow (i) \) trivially.

\[ \square \]

**Proposition 2.2** The set \( C \) has finitely many connected components, each of which is closed and on each of which \( H \) is constant. Furthermore, if all the principal minors of \( R \) are invertible, then \( C \) consists of at most \( 2^d - 1 \) points.

Proof: By (3) (ii), \( C \) is the union over all \( 2^d - 1 \) faces \( F \) of the sets \( C_F = \{ v : D_vH|_F(v) = 0 \} \). By (4) and the comment following, \( D_vH|_F \) is linear, so \( C_F \) is a closed, convex, connected set. It is easy to see that \( H \) is constant on \( C_F \) by integrating \( D_vH|_F \). The first part of the proposition follows since each connected component of \( C \) is the union of some of the \( C_F \). For the second part, fix a face \( F \) and let \( R_F \) be the matrix gotten from \( R \) by deleting rows and columns indexed by those \( i \) for which \( v_i = 0 \) for all \( v \in F \). If this is invertible, then equation (3) (iii) implies that the only possible element of \( C \) in the interior of \( F \) is whichever multiple of \( (1, \ldots, 1)R_F^{-1} \) lies on the unit simplex.

\[ \square \]

If all the off-diagonal entries of \( R \) are positive, it is immediate that \( M(v) \) is irreducible for all \( v \in \triangle \). Conversely, if \( R_{ij} = 0 \) for some \( i \neq j \), then \( M(v) \) is reducible when \( v \) is
any nontrivial combination of $e_i$ and $e_j$. Thus it a necessary and sufficient condition for $C_0$ to be empty is that $R_{ij} > 0$ off of the diagonal. In any event, $C_0$ is a union of proper faces of $\triangle$. The following corollary to Theorem 1.1 is now immediate.

**Corollary 2.3** If all the off-diagonal entries of $R$ are positive and all the principal minors of $R$ are invertible, then $V(n)$ converges almost surely.

\[
\square
\]

3 Proofs of convergence results

The proof of Theorem 1.1 begins with a lemma giving a lower bound on the expected growth of $H(V(n))$ when $V(n)$ is not near $C \cup C_0$.

**Lemma 3.1** Let $N$ be a closed subset of the simplex, with $N \cap (C \cup C_0) = \emptyset$. Then there exist an $N$, $L$ and $c > 0$ such that for any $n > N$, $E(H(V(n + L))|V(n)) > H(V(n)) + c/n$ whenever $V(n) \in N$.

Proof: For any $n$, let $M_n(n), M_n(n + 1), \ldots$ denote a Markov chain beginning at $Y_n$ at time $n$, whose transition matrix thereafter does not change with time and is given by $M(V(n))$. Let $S'(n) = S(n)$ and for $i > n$, let $S'(i) = S'(i - 1) + e_{M_n(i)}$, where $e_j$ is the $j^{th}$ standard basis vector. Let $V'(i) = S'(i)/i$.

First I claim that the lemma is true with the Markov process $V'$ substituted for $V$. By Lemma 2.1, $D_y H(\pi(v) - v)$ is nonzero on $N$, so by compactness it is bounded below by some $c_0$ on $N$. Choose any $c_1 < c_0$. The occupation measure of a process between times $N$ and $N + L$ can change by at most $L/(N + L)$ in total variation.
Since $H$ is smooth, it is possible to choose $N/L$ large enough so that whenever $n \geq N$, $H[V'(n) + (L/(N + L))\pi(V(n)) - V(n)] > c_1L/(n + L)$. By the Markov property, $(S'(n + L) - S(n))/L$ approaches a point-mass at $\pi(v)$ in distribution as $L$ increases. In fact, the rate of convergence of $M^k(V(n))w$ to $\pi(V(n))$ is exponential and controlled by the second-largest eigenvalue of $M(V(n))$ according to the Perron-Frobenius theorem. If $M(V(n))$ is aperiodic, then since $M(v)$ varies continuously with $v$, eigenvalues are continuous, and the non-degeneracy hypothesis says that $N$ contains no points where the second-largest eigenvalue is 1, the second-largest eigenvalue is bounded away from 1. It follows that a large enough $L$ may be chosen uniformly in $v$ so that $E[H(V'(n + L)) - H(V(n))| F_n] > c/n$ for any $c_2 < c_1$, and the claim is established. If $M(V(n))$ is periodic, then it has period 2 and a simple eigenvalue at $-1$; the claim follows in this case from grouping together pairs of times $2n$ and $2n + 1$.

Now couple the Markov chain $V'(n + i)$ to $V(n + i)$ in such a way so the two move identically for as long as possible. Formally, define $\{M_n(i)\}$ and $\{Y_i\}$ on a common measure space so that if $Y_j = M_n(j)$ for all $n < j < n + k$ then

$$P(Y_{n+k} \neq M_n(n + k)) | Y_{n+k-1} = i = \sum_j \frac{1}{2} |M_{ij}(V(n + k)) - M_{ij}(V(n))|.$$ 

Picking $c < c_2$ and $N/L$ large enough so that

$$(L^2/N)(L/N)\|DH\|_{op} < (c_2 - c)/N,$$ 

(5)

the coordinates of $V$ cannot change by more than $L/N$ in $L$ steps, so the probability of an uncoupling at any of the $L$ steps is bounded by $L^2/N$. Then $E|H(V(n + L)) - H(V'(n + L))| < (c_2 - c)/N$ by (5), and combining this with the earlier claim proves the lemma. 

Before proving Theorem 1.1, here is a sketch of the argument. On any set $\mathcal{N}$ away from $C \cup \mathcal{C}_0$, Lemma 3.1 says the expected value of $H(v(n))$ grows, provided you sample at time intervals of size $L$. The cumulative differences between $H(v(n + L))$ and $E(H(v(n + L))$
form a convergent martingale, so $H(V(n))$ itself is growing at rate $c/n$ when $V(n) \in \mathcal{N}$. The rate of change in position of $V(n)$ is also order $1/n$ per step, so if $V$ goes from one given point of $\mathcal{N}$ to another, $H(V(n))$ increases by an amount independent of time. The only way it can decrease again is for $V(n)$ to leave $\mathcal{N}$ at a place where $H$ is large and re-enter where $H$ is small. The effect of such a possibility can be made arbitrarily small because $H$ is nearly constant on the connected components of $\triangle \setminus \mathcal{N}$.

Proof: of Theorem 1.1: Since the connected components, $C_1, \ldots, C_k$ of $C \cup C_0$ are closed, $m = \min\{d(C_i, C_j)\} > 0$. Pick any $r < m/3$. Let

$$N_1^i = \{v : d(v, C_i) < r\}$$
$$\mathcal{N}_1 = \triangle \setminus \bigcup_{i=1}^k N_1^i. \quad (6)$$

Note that

$$i \neq j \Rightarrow d(N_1^i, N_1^j) > r. \quad (7)$$

By the preceding lemma with $\mathcal{N} = \mathcal{N}_1$, $c_1, L_1, N_1$ can be found for which $n \geq N_1$ implies $E(H(V(n + L)) \mid V(n)) \geq H(V(n)) + c/n$. Pick any $L' > L_1$ and define

$$N_2^i = N_1^i \cap \{v : |H(v) - H(C_i)| < rc/2L'\}$$
$$\mathcal{N}_2 = \triangle \setminus \bigcup_{i=1}^k N_2^i. \quad (8)$$

Figure 1 gives an example of these definitions when $d = 3$; the heavy lines are the boundary of $\mathcal{N}_1$ and the lighter lines are the boundary of $\mathcal{N}_2$.

Apply the lemma to $\mathcal{N}_2$ to get $N_2, c_2$ and $L_2$. Define the process $\{U(n)\}$ that samples $V(n)$ at intervals of $L_1$ on $\mathcal{N}_1$ and $L_2$ elsewhere, by

$$U(n, \omega) = V(f(n, \omega))$$

where
\[ f(1, \omega) = \max\{N_1, N_2\} \] and
\[
f(n + 1, \omega) = \begin{cases} f(n, \omega) + L_1 & \text{if } V(f(n, \omega)) \in N_1; \\ f(n, \omega) + L_2 & \text{if } V(f(n, \omega)) \notin N_1. \end{cases}
\]

Clearly, \(U(n)\) converges if and only if \(V(n)\) converges. Letting \(U(n) = H(U(n))\), write \(U(n) = M(n) + A(n)\) where \(\{M(n)\}\) is a martingale and \(\{A(n)\}\) is a predictable process with respect to \(\mathcal{F}_{f(n)}\). The key properties needed are

\begin{align*}
M(n) & \text{ converges almost surely} \quad (8) \\
A(n + 1) & \geq A(n) + c/n \text{ if } U(n) \in N_1 \quad (9) \\
A(n + 1) & \geq A(n) \text{ if } U(n) \in N_2. \quad (10)
\end{align*}

To verify (8), note that \(|U(n + 1) - U(n)| \leq \max\{L_1, L_2\}/f(n) = O(1/n)\), since by (4), \(H\) is Lipschitz on \(\triangle\). Then \(|M(n + 1) - M(n)| = O(1/n)\) as well, so \(M(n)\) converges in \(L^2\), hence almost surely. Properties (9) and (10) are evident from the construction.

The next thing to show is Claim 1: \(U(n) \in N_2^a\) infinitely often for at most one \(a\) almost surely. Consider any sample path \(U(1), U(2), \ldots\) For \(n < t\), define the event \(B(a, b, n, t, \omega)\) to occur if

\[ U(n) \in N_2^a \text{ and } U(t) \in N_2^b \text{ with } U(i) \in N_2 \text{ for all } i \text{ such that } n < i < t. \quad (11) \]

If \(B(a, b, n, t, \omega)\) occurs, let

\[
r = \max\{i : n \leq i < t \text{ and } U(i) \in N_1^a\} \quad \text{and} \quad s = \min\{i : n \leq i < t \text{ and } U(i) \in N_1^b\}
\]

be respectively the last exit time of \(N_1^a\) and the first entrance time of \(N_1^b\). The dotted path in figure 1 gives an example of this. By (9) and (10),

\[ A(i + 1) - A(i) \geq c/i \text{ for } r < i < s \]
\[ A(i + 1) - A(i) \geq 0 \text{ for } n < i < t. \]

Then

\[
\begin{align*}
A(t) - A(n) & = [A(t) - A(s)] + [A(s) - A(r + 1)] \\
& + [A(r + 1) - A(n + 1)] + [A(n + 1) - A(n)] \\
& \geq 0 + \left( \sum_{i=r+1}^{s-1} c/i \right) + 0 - L_2/n \\
& = O(1/n) + (c/L_1) \sum_{i=r}^{s-1} L_1/i \\
& \geq O(1/n) + (c/L_1) \sum_{i=r}^{s-1} |U(i + 1) - U(i)| \\
& \geq O(1/n) + (c/L_1)|U(s) - U(r)| \\
& > O(1/n) + rc/L_1
\end{align*}
\]

by (7). Now \( U(t) - U(n) \leq H(C_b) - H(C_a) + rc/L' \) by the construction of \( N_2. \) So \( M(t) - M(n) \leq H(C_b) - H(C_a) + rc/L' - rc/L_1 + O(1/n). \) If \( H(C_b) \leq H(C_a), \) the choice of \( r \) guarantees that this expression is strictly negative and bounded away from 0 for large \( n. \) Therefore if \( M(n)(\omega) \) converges, then \( B(a, b, n, t, \omega) \) happens only finitely often for \( a, b \) such that \( H(C_b) \leq H(C_a). \) But then it happens only finitely often for any \( a \neq b, \) since \( U \) can make only \( k - 1 \) successive transitions from \( N_2^a \) to \( N_2^b \) with \( H(C_b) > H(C_a). \) Thus the almost sure convergence of \( M(n) \) implies that \( U(n) \in N_2^a \) infinitely often for at most one \( a \) almost surely and Claim 1 is shown.

In other words, transitions between small neighborhoods of \( \mathcal{C}_i \) and \( \mathcal{C}_j \) eventually cease for \( i \neq j. \) Claim 2 is that \( V(n) \) may not oscillate between a small neighborhood of \( \mathcal{C}_i \) and a set bounded away from \( \mathcal{C}. \) To show this, require now that \( r < m/6. \) With \( N_1 \) and \( N_2 \) defined as before, define \( N_3 \subseteq N_1 \) by (6) with \( 2r \) in place of \( r. \) Since \( 2r < m/3, \) equation (7) holds with \( N_3 \) in place of \( N_1. \) An argument identical to the one establishing
Claim 1 now shows that with probability 1 there are only finitely many values of n and t for which

\[ U(n) \in \mathcal{N}_2^a, U(i) \in \mathcal{N}_3 \text{ and } U(t) \in \mathcal{N}_2^a \text{ for } n < i < t. \]

[The argument again: A(i) is nondecreasing when U(n) ∈ \mathcal{N}_1^a and increases by at least the fixed amount rc/L each time U makes the transit from \mathcal{N}_1^a to \mathcal{N}_3. The increase in A is greater than the greatest difference in values of H taken at two points of \mathcal{N}_2^a, so the martingale M must change by at least rc/L - rc/L' during every transit. Since M converges, this happens finitely often.]

Claim 3 is that the event \{ω : U(t, ω) ∈ \mathcal{N}_1 for all t > n\} has probability 0 for each n; it is proved in an identical manner. Putting together Claims 1 and 3, it follows that for any small r there is precisely one a for which U(n) ∈ \mathcal{N}_1^a infinitely often. Then by Claim 2 for a different r, \mathcal{N}_3 stops being visited, so letting r → 0 proves the theorem. □

The proof of Theorem 1.4 is just an easier version of the proof of Theorem 1.1.

Sketch of proof of Theorem 1.4: A process U(n) may be defined as in the previous proof, so that V(n) converges iff U(n) converges and so that U(n) \triangleq H(U(N)) breaks into a martingale M(n) and a predictable process A(n). Note that the argument showing an \( L^2 \) bound of c/n on M(∞) - M(n) still works conditionally on U(n). By a standard maximal inequality, given any \( \epsilon > 0 \), an n may be chosen large enough so that \( \mathbb{P}(\inf\{M(n) - M(n+i) : i > 0\} < -\epsilon | U(n)) < \epsilon \). The assumptions of the theorem imply the existence of an \( \epsilon \) for which the component B of \( H^{-1}[a - 2\epsilon, a] \) is disjoint from \((\mathcal{C} \cup \mathcal{C}_0) \setminus A\), where a is the value of H on A. Now for sufficiently large n, the event \( U(n) \in H^{-1}[a - \epsilon, a] \cap B \) has positive probability. Conditional on this event, the probability that \( M(n+i) - M(n) \) never goes below \( -\epsilon \) has been shown to be less than \( \epsilon \) for large n. Since \( \text{dist}(U(n), \mathcal{C} \cup \mathcal{C}_0) \rightarrow 0 \) by Theorem 1.1, and U(n) cannot leave B without U(n) becoming less than \( a - 2\epsilon \), it follows that \( \text{dist}(U(n), A) \rightarrow 0 \), proving the theorem. □
To prove Theorem 1.3, begin by seeing why it should be true. With $p$ as in the statement of the theorem, equation (3) (iii) says that the $N_i$ have a common value, $\lambda$, for those $i$ such that $p_i > 0$. Assuming (2) for a given $e_k$ and using equation (4) for $DH$ shows that $N_k > N_j = \lambda$. So

$$\sum_i R_{ki}p_i/N_i = \sum_{p_i > 0} R_{ki}p_i/\lambda = N_k/\lambda = 1 + b$$

(12)

for some $b > 0$, $k$ such that $p_k = 0$. Now when $V(n)$ is close to $p$, $v_k(n)$ will be close to but not equal to zero. The expected number of visits to state $k$ during a period of time from $n$ to $n + T$ in which the occupation measure is close to $p$ will be approximately $T \sum_i p_i(R_{ik}v_k/N_i) = T v_k N_k/\lambda = (1 + b)Tv_k$. In other words, $v_k$ will begin to increase and $p$ should be an unstable point with no possibility of $V(n)$ converging there. The actual proof will consist of making this rigorous.

To avoid bogging down in trivialities, $S(n)$ and $V(n)$ will be used to stand for $S(\lfloor n \rfloor)$ and $V(\lfloor n \rfloor)$. Inequalities will be verified as if $n$ were an integer; it is always possible to choose epsilons and deltas a little bit smaller to compensate for the roundoff errors. Begin by recording a few propositions whose proofs are omitted when elementary.

**Proposition 4.1** Fix $p$ and let $N_1$ be a neighborhood of $p$. For any $\delta > 0$ there is a neighborhood $N$ of $p$ included in $N_1$ such that for all $n > 1/\delta$, the two conditions

(i) $V(n) \in N$ and

(ii) $V(n + \delta n) \in N$

imply

(iii) $(S(n + \delta n) - S(n))/\delta n \in N_1$.

$\square$
Proposition 4.2 Let \( p, k, b \) be such that (12) holds and let \( S \) be any vector function of \( n \). Then there is an \( \epsilon > 0 \) and a neighborhood \( \mathcal{N}_1 = \{ v \in \triangle : |v - p| < \epsilon \} \) such that for all \( \delta > 0 \) and for all \( n \), the conditions \( V(n) \in \mathcal{N}_1 \) and \( (S_i(n + \delta n) - S_i(n))/\delta n \geq p_i - \epsilon \) for all \( i \) imply
\[
\sum_i (S_i(n + \delta n) - S_i(n))R_{ik}v_k(n)/(1 + \delta)N_i(n) > \frac{1 + b/2}{1 + \delta}S_k(n). \tag{13}
\]
Proof: As \( \epsilon \to 0 \), \( 1/n \) times the left-hand side converges to \( \delta p_k(n)/(1 + \delta) \sum_i p_i(n)R_{ik}/N_i(n) \) = \( \delta p_k(n)(1 + b)/(1 + \delta) \) while \( 1/n \) times the right-hand side converges to \( \delta p_k(n)(1 + b/2)/(1 + \delta) \). Since the convergence is uniform in \( \delta \), the result follows. \( \square \)

Proposition 4.3 Let \( b > 0 \) and \( \epsilon_1 > 0 \) be given. Let \( \{ B_\alpha \} \) be a collection of independent Bernoulli random variables with \( \mathbb{E}(\sum_\alpha B_\alpha) \geq (1 + b)L \). There exists an \( L_0 \) such that whenever \( L > L_0 \), \( \mathbb{P}(\sum_\alpha B_\alpha/L > 1 + b/2) > 1 - \epsilon_1 \). \( \square \)

Proof of Theorem 1.3: By hypothesis, condition (2) holds, and hence (12) holds for some choice of \( p, k \) and \( b \) which are fixed hereafter. Pick \( \epsilon \) and \( \mathcal{N}_1 \) according to Proposition 4.2. Apply Proposition 4.1 to \( \mathcal{N}_1 \) and \( p \) with \( \delta = 1 \wedge (1 + b/2)/(1 + b/4) - 1 \) to obtain a neighborhood \( \mathcal{N} \) of \( p \) with the appropriate properties. Temporarily fixing \( n \), define the event \( B_n \) by \( V(i) \in \mathcal{N} \) for all \( n \leq i \leq (1 + \delta)n \). Define stopping times \( \{ \tau_{i,r} \} \) and a family of Bernoulli random variables \( \{ B_{i,r} \} \) as follows.

Let \( \tau_{i,r} \leq \infty \) be the \( r \)th time after \( n \) that \( Y_j = i \), so formally \( \tau_{i,0} = n \) and \( \tau_{i,r+1} = \inf\{ j > \tau_{i,r} : Y_j = i \} \). Let \( B_{i,r} \) be independent and Bernoulli with
\[
\mathbb{P}(B_{i,r} = 1) = R_{ki}v_k(n)/(1 + \delta)N_i(n) \tag{14}
\]
as coupled to the variables \( \{ Y_i \} \) so that if \( B_{i,r} = 1 \) and \( \tau_{i,r} \leq (1 + \delta)n \) then \( Y_{1+\tau_{i,r}} = k \).
To verify that this construction is possible, check that the probability of a transition from vertex $i$ to vertex $k$ never drops below the quantity in (14):

$$
P(Y_{1+\tau_{i,r}} = k \mid \mathcal{F}_{\tau_{i,r}}) \geq \frac{(n/\tau_{i,r})R_{ki}v_k(n)/N_i(n)}{(1/(1+\delta)) R_{ki}v_k(n)/N_i(n)}$$

for $\tau_{i,r} < (1+\delta)n$.

Now consider the subcollection $A \overset{\text{def}}{=} \{(i, r) : r \leq \delta n(p_i - \epsilon)\}$. By Proposition 4.1, $\tau_{i,r} \leq (1+\delta)n$ whenever the event $\mathcal{B}_n$ holds. Meanwhile,

$$\mathbb{E}(\sum_{\alpha \in A} B_\alpha) = \sum_i \delta n(p_i - \epsilon)R_{ki}v_k(n)/(1+\delta)N_i(n).$$

By Proposition 4.2, this quantity is at least $\delta(1 + b/2)S_k(n)/(1 + \delta)$ which is at least $\delta(1 + b/4)S_k(n)$ by choice of $\delta$. Apply Proposition 4.3 to the collection $\{B_\alpha : \alpha \in A\}$, with $b$ replaced by $b/4$ and $\epsilon_1$ to be chosen later to obtain a value for $L_0$. Now calculate the conditional expectation $\mathbb{E}(\ln(v_k((1+\delta)n)) \mid \mathcal{F}_n, S_k(n) > L_0)$. By Proposition 4.3 and the coupling,

$$\mathbb{P}(\mathcal{B}_n \text{ and } S_k((1+\delta)n) - S_k(n) \geq \delta(1 + b/8)S_k(n) \mid \mathcal{F}_n, S_k(n) > L_0)$$

$$> \mathbb{P}(\mathcal{B}_n \mid \mathcal{F}_n, S_k(n) > L_0) - \epsilon_1.$$ 

When $S_k((1+\delta)n) - S_k(n) \geq \delta(1 + b/8)S_k(n)$, it follows that $v_k((1+\delta)n) \geq v_k(n)(1 + b/8(1+\delta)) \geq v_k(n)(1 + b/16)$. Therefore

$$\mathbb{E}(\ln(v_k((1+\delta)n)) \mid \mathcal{F}_n, S_k(n) > L_0)$$

$$\geq (\mathbb{P}(\mathcal{B}_n \mid \mathcal{F}_n, S_k(n) > L_0) - \epsilon_1) \ln((1 + b/16)v_k(n))$$

$$+ (1 - \mathbb{P}(\mathcal{B}_n \mid \mathcal{F}_n, S_k(n) > L_0) + \epsilon_1) \ln(v_k(n)/(1 + \delta))$$

$$\geq \ln(v_k(n)) + \ln(1 + b/32) - KP(\mathcal{B}_n^c \mid \mathcal{F}_n, S_k(n) > L_0)$$

(15)
for $K \ln(1 + \delta)(1 + b/16)$, when $\epsilon_1$ is sufficiently small. To conclude from this that $P(\mathbf{V}(n) \to \mathbf{p}$ and $S_k(n) > L_0$ for some $n) = 0$, write $T(n) = (1 + \delta)^n L_0$, $\mathcal{G}_n = \mathcal{F}_{T(n)}$, $X_n = \ln(v_k(T(n)))$, $\beta = c/2K$, $T = \inf\{n : \mathbf{V}(i) \notin \mathcal{N}$ for some $L_0 \leq i \leq T(n)\}$ and calculate

$$E \mathbf{X}_{n \wedge T} = X_0 + \sum_{i=0}^{n-1} E(1_{T > i}(X_{i+1} - X_i) | \mathcal{G}_i)$$

$$\geq X_0 + \sum_{i=0}^{n-1} E 1_{T > i} \left[ (c - K\beta) \cdot 1_{P(T = i+1 | \mathcal{G}_n) \leq \beta} - K \cdot 1_{P(T = i+1 | \mathcal{G}_n) > \beta} \right]$$

by equation 16

$$\geq X_0 + \sum_{i=0}^{n-1} E 1_{T > i} \left[ (c - K\beta) \cdot (1 - \beta^{-1}P(T = i+1 | \mathcal{G}_n, T > i)) - K \cdot \beta^{-1}P(T = i+1 | \mathcal{G}_n, T > i) \right]$$

$$\geq X_0 + \sum_{i=0}^{n-1} (c - K\beta)1_{T > i} - \beta^{-1}(c + K - K\beta)P(T = i + 1)$$

$$\geq X_0 + n(c - K\beta)P(T > n) - \beta^{-1}(c + K - K\beta).$$

Since $c - K\beta$ was chosen to be positive, $P(T > n)$ must go to zero, showing that $\mathbf{V}(n) \to \mathbf{p}$ and $S_k(n) > L_0$ eventually is impossible.

Finally, to show that $P(\mathbf{V}(n) \to \mathbf{p}$ and $S_k(n) \leq L_0$ for all $n) = 0$, note that since $N_k(\mathbf{p}) > 0$, there is a sufficiently small neighborhood $\mathcal{N}$ of $\mathbf{p}$ for which $P(Y_{i+1} = k | \mathcal{F}_i, \mathbf{V}(n) \in \mathcal{N})$ is always at least a constant times $n^{-1}$. Borel-Cantelli implies that $k$ is visited infinitely often whenever $\mathbf{V}(n)$ remains in $\mathcal{N}$, and this finishes the proof of Theorem 1.3. □
5 Proof of Theorem 1.2

Begin with a proof of the equivalences:
\[ p \text{ fails to be a maximum for } H \]
\[ \iff R \text{ has more than one positive eigenvalue} \]
\[ \iff D_p(\pi - I) \text{ has a positive eigenvalue} \]

The matrix $R$ can be viewed as a symmetric bilinear form whose quadratic form gives $H$ when restricted to $\triangle$. Let $W = \triangle - p$ be the translation of $\triangle$ containing the origin. For $w \in W$,
\[ R(w, p) = w^T R p = w \cdot \lambda \cdot (1, \ldots, 1) = 0 \]
where $\lambda$ is the common value of the $N_i$. Then
\[ R(w + cp, w + cp) = R(w, w) + R(cp, cp) = R|_W(w) + c^2 \lambda \] (17)
so the quadratic form $R(v, v)$ decomposes into the sum of $R|_W$ and a positive form on the one-dimensional subspace spanned by $p$. Then $R$ has precisely one more positive eigenvalue than the quadratic form $R|_W$. But equation (17) with $w = v - p$ shows that $H(v) = R|_W(v - p) + \lambda$ so $H$ has a strict maximum at $p$ if and only if $R|_W$ has a strict maximum at the origin. Since $R$ has no zero eigenvalues, $R|_W$ will have a strict maximum when it has a maximum, which happens when it has no positive eigenvalues.

For the second equivalence, note that $\pi$ is smooth on the interior of $\triangle$, so $D_p(\pi - I)$ exists. Let $\tilde{T}$ be the operator on $\mathbb{R}^d$ whose matrix in the standard basis is given by
\[ \tilde{T}_{ij} = R_{ij}p_i/\lambda. \]
I claim that $\tilde{T} = D_p(\pi - I)$ on $W$. Indeed, using Definition (3) to define $\pi$ on all of $\mathbb{R}^d$ and differentiating shows that the matrix representation for $D_p(\pi - I)$ is given by
\[ [D_p(\pi - I)]_{ij} = \left. \frac{\partial}{\partial e_j} (\pi(v))_i \right|_{v=p} - \delta_{ij} \]
\[
\begin{align*}
\frac{\partial}{\partial e_j} v_i N_i H(v) \bigg|_{v=p} &= -\delta_{ij} \\
\left( \frac{R_{ij} v_i}{H} + \frac{\delta_{ij} N_i}{H} - \frac{v_i N_i \partial H / \partial e_j}{H^2} \right) \bigg|_{p} - \delta_{ij} &= R_{ij} p_i / \lambda - 2 p_i
\end{align*}
\]

(using the fact that all the \(N_i\) have a common value \(\lambda = H(p)\) and the identity \(\frac{\partial H}{\partial e_j} = 2 N_j\)).

Then the matrices for \(\tilde{T}\) and \(D_p(\pi - I)\) differ by a matrix with constant rows, hence define the same operator on \(W\). Now let \(\text{diag}(p)\) be the diagonal matrix with \(i, i\) entry equal to \(p_i\) and observe that \(\tilde{T} = \text{diag}(p)R/\lambda\). Since \(R\) is symmetric and \(\text{diag}(p)\) is positive definite, \(\tilde{T}\) must be diagonalizable with real eigenvalues and has the same signature as \(R\) (see [4, Theorem 6.23 and 6.24 page 232]). Since \(\tilde{T}\) has \(p\) as a positive eigenvalue and \(W\) as an invariant subspace, it has one more positive eigenvalue than \(D_p(\pi - I)\) and the conclusion follows.

To finish proving Theorem 1.2, it remains to show that \(v(n)\) cannot converge to an interior point where \(D_p(\pi - I)\) has a positive eigenvalue. The method of proof is from [7], the first step being a construction of a scalar function which measures “distance from \(p\) in an unstable direction” ([7, Proposition 3]).

**Lemma 5.1** Under the assumptions of Theorem 1.2, suppose that \(D_p(\pi - I)\) has a positive eigenvalue. Then there is a function \(\eta\) from a neighborhood of \(p\) to \([0, \infty)\) such that \(D_v \eta(\pi(v) - v) \geq k_1 \eta(v)\) in a neighborhood of \(p\) for a constant \(k_1 > 0\). Furthermore, \(\eta\) is the square root of a smooth function (whose gradient necessarily vanishes whenever the function vanishes) but whose second partials are not all zero (thus \(\eta\) is not differentiable where it vanishes). It follows from this that \(\eta\) is Lipschitz and that \(\eta(v + w) \geq \eta(v) + D_v \eta(w) + k_2 \|w\|^2\) in a neighborhood of \(p\), where \(D_v \eta(F)\) may be any of the support hyperplanes to the graph of \(\eta\) at points where \(\eta\) vanishes.

Use this lemma and a sequence of appropriately chosen stopping times to convert
questions about convergence of \( V(n) \) into questions about the convergence of a scalar stochastic process. To do this, fix a neighborhood \( \mathcal{N} \) of \( p \) in which all coordinates are bounded away from zero. Let \( L(v) \) be the mean recurrence time to state 1 for the Markov chain \( M(v) \) and let \( L_{\text{max}} = \sup_{v \in \mathcal{N}} L(v) \). Pick \( N_0 > 2L_{\text{max}} \) and define \( \sigma_0 = \inf\{k \geq N_0 : Y_k = 1\} \) and \( \sigma_{n+1} = \inf\{k > \sigma_n : Y_k = 1\} \) to be the successive hitting times for state 1. Let \( \tau = \inf\{k \geq N_0 : V(k) \notin \mathcal{N}\} \) and let \( \tau_i = \tau \land \sigma_i \). For the remainder of the section, let \( \mathbb{E} \) and \( \mathbb{P} \) denote conditional expectation and conditional probability with respect to \( \mathcal{F}_{\tau_n} \). The following facts are elementary.

**Proposition 5.2**

(i) The distribution of \( \tau_{n+1} - \tau_n \) has finite conditional expectation and variance. Specifically,

\[
\mathbb{P}(\tau_{n+1} - \tau_n \geq k + 1) < e^{-\alpha k}
\]

for some \( \alpha > 0 \).

(ii) For any \( \epsilon > 0, N_0 \), there is a constant \( c_1 \) such that

\[
\mathbb{P}(n \leq \tau_n \leq c_1 n \text{ for all } \tau_n \leq \tau) > 1 - \epsilon.
\]

(iii) For any \( \epsilon > 0, \gamma < 1, N_0 \) and \( r \) may be chosen large enough so that

\[
\mathbb{P}(\tau_{n+1} - \tau_n \geq n^{1-\gamma} \text{ for some } n \geq r) < \epsilon.
\]

Let \( U(n) = V(\tau_n) \), let \( S_n = \eta(U(n)) \) and let \( X_n = S_n - S_{n-1} \). The following estimate shows that the expected increment in \( U \) from time \( n \) to \( n + 1 \) is close to the value given by the Markov approximation.
Proposition 5.3 For any $n > 0$,

$$
\left| \mathbb{E}(U(n+1) - U(n)) - L(U(n)) \frac{\pi(U(n)) - U(n)}{\tau_n} \right| = O(\tau_n^{-2}).
$$  \tag{18}

Proof: Couple the process $\{Y_i : i \geq \tau_n\}$ to a Markov chain $Y'_n$ with $Y_{\tau_n} = 1$ and transition matrix $M(U(n))$ in such a way that the two processes remain identical for as long as possible. Define $V'_n, S'_n, \tau'_n$ and $U'_n$ analogously to the unprimed variables. Establish first that

$$
\mathbb{E}|U(n + 1) - U'(n + 1)| = O(\tau_n^{-2}).
$$  \tag{19}

To see this, observe that since transition probabilities for $Y$ and $Y'$ differ by at most $k/\tau_n$ at time $\tau_n + k$, the conditional probability of the two processes uncoupling before time $\tau_{n+1}$ is at most

$$
\sum_{k \geq 0} P(\tau_{n+1} - \tau_n > k) k/\tau_n \leq e^{-\alpha}/(1 - e^{-\alpha})^2 \tau_n
$$  \tag{20}

according to Proposition 5.2 (i). On the other hand, $\mathbb{E}(\tau'_{n+1} - (\tau_n + k) \mid \mathcal{F}_{k+\tau_n})$ and $\mathbb{E}(\tau'_{n+1} - (\tau_n + k) \mid \mathcal{F}_{k+\tau_n})$ are bounded by $L_{\text{max}}$ and $(1 - e^{-\alpha})^{-1}$ respectively on the event of the uncoupling occurring at time $k + \tau_n$, which implies that

$$
\mathbb{E}(|v(\tau_{n+1}) - \tilde{v}(\bar{\tau}_{n+1})| \mid \text{uncoupling before } \tau_{n+1})
\leq \sup_k \mathbb{E}(|v(\tau_{n+1}) - v(\tau_n)| + |\tilde{v}(\bar{\tau}_{n+1}) - v(\tau_n)| \mid \text{uncoupling occurs at } \tau_n + k)
\leq \sup_k (1/\tau_n) (\mathbb{E}(\tau_{n+1} - \tau_n - k + 1
+ \bar{\tau}_{n+1} - \tau_n - k + 1) \mid \text{uncoupling occurs at } \tau_n + k)
\leq (1/\tau_n)(L_{\text{max}} + 1/(1 - e^{-\alpha}) + 2).
$$  \tag{21}

Combining (20) and (21) gives (19).
The quantity $U(n + 1) - U(n)$ in the LHS of equation (18) may now be replaced by the quantity $U'(n + 1) - U(n)$, since the two are within $O(\tau_n^{-2})$ in expectation. Since $Y'$ is a Markov chain, the following identity holds:
\[
\mathbb{E}(S(\tau_n) - S(n)) = L(U(n))\pi(U(n)).
\] (22)

Component by component, we then have
\[
\mathbb{E}(U'_i(n + 1) - U'_i(n)) = \mathbb{E}(S'_i(\tau_{n+1})/\tau_{n+1} - S_i(\tau_n)/\tau_n)
\]
\[
= \mathbb{E}\left(\frac{1}{\tau_n} (S'_i(\tau_{n+1}) - S_i(\tau_n) - (\tau_{n+1} - \tau_n)S_i(\tau_n)/\tau_n)\right) - Q
\]
\[
= \frac{1}{\tau_n} L(U(n))\pi_i(U(n)) - U_i(n) - Q
\]
according to (22), where
\[
Q = \mathbb{E}\left(\frac{\tau_{n+1} - \tau_n}{\tau_{n+1}} \cdot \frac{S'_i(\tau_{n+1}) - S_i(\tau_n) - (\tau_{n+1} - \tau_n)S_i(\tau_n)/\tau_n}{\tau_n}\right).
\]
The denominator of $Q$ is at least $\tau_n^2$ and the numerator is bounded by the product of two geometric random variables according to Proposition 5.2, so $|Q| = O(\tau_n^{-2})$ and the proposition is proved.
\[\Box\]

Use this estimate to prove the following proposition, which together with the Lemma 5.5 proves Theorem 1.2.

**Proposition 5.4** Let $S_n$ and $X_n$ be defined from $V$ as above. Let $N$ remain fixed as in the paragraph before Proposition 5.2. For any $\epsilon > 0$ there are constants $b_1, b_2, c > 0$ and $\gamma > 1/2$ and an $N$ such that whenever $N_0 > N$ then
\[
P(B \mid F_{N_0}) > 1 - \epsilon,
\] (23)
where \( B \) is the event that either equations (24) - (27) are satisfied for all \( n > N_0 \) or else \( V(n) \) at some point leaves \( \mathcal{N} \).

\[
\begin{align*}
\mathbb{E}(X_{n+1}^2 + 2X_{n+1}S_n) &\geq b_1/n^2 \quad (24) \\
\mathbb{E}(X_{n+1}S_n1_{S_n > c/n}) &\geq 0 \quad (25) \\
\mathbb{P}(|X_{n+1}| \leq 1/(n+1)^\gamma) &= 1 \quad (26) \\
\mathbb{E}(X_{n+1}^2) &\leq b_2/n^2 \quad (27)
\end{align*}
\]

**Lemma 5.5** If (23) holds for a nonnegative stochastic process \( S_n = S_0 + \sum_{i=1}^n X_i \), then \( \mathbb{P}(S_n \to 0) = 0 \).

Lemma 5.5 is a variant on an argument from [7], whose proof can be outlined as follows.

First assume that (23) holds with \( \epsilon = 0 \), i.e. that (24) - (27) hold almost surely, and show in the following three steps (A)-(C) that \( \mathbb{P}(S_n \to 0) = 0 \). Let \( k \) be any positive real number less than \( \sqrt{b_1/2} \) and without loss of generality restrict \( n \) to be at least \( 4c^2/k \) so that \( k/2\sqrt{n} > c/n \) and \( c \) is the constant in condition (25).

(A) Claim: given any \( S_n \), the probability of finding \( S_M > k/\sqrt{n} \) for some \( M \geq n \) is at least \( 1/2 \).

Proof: Assume without loss of generality that \( S_n < k/\sqrt{n} \). Let \( \sigma \) be the first \( i \geq n \) for which \( S_i > k/\sqrt{n} \). Then for any \( M > n \),

\[
\begin{align*}
\mathbb{E}(S_{\sigma \wedge M}^2) &\geq S_n^2 + \sum_{i=n}^{M-1} \mathbb{E}(S_{\sigma \wedge (i+1)}^2 - S_{\sigma \wedge i}^2) \\
&\geq S_n^2 + \sum_{i=n}^{M-1} \mathbb{E}(S_{\sigma \wedge (i+1)}^2 - S_{\sigma \wedge i}^2)
\end{align*}
\]

22
\[ S_n^2 + \sum_{i=n}^{M-1} \mathbf{E}(1_{s \geq i}(X_{i+1}^2 + 2X_{i+1}S_i)) \]

\[ \geq \mathbf{P}(s > M) \sum_{i=n}^{M-1} b_1/i^2 \]

by (25)

\[ \geq \mathbf{P}(s > M)b_1/n. \]

But condition (26) implies that \( S_{\sigma \wedge i} \) never gets much more than \( k/\sqrt{n} \) and since \( k^2 < 2b_1 \) this forces \( \mathbf{P}(s > M) < 1/2 \) and the claim is proved.

(B) Claim: given that \( S_n > k/\sqrt{n} \) the probability that \( S_M \) will never return to the interval \( x < k/\sqrt{n} \) for \( M > n \) is at least \( a = 4b_2/(4b_2 + k^2) \).

Proof: Assume \( S_n > k/\sqrt{n} \). Let \( \sigma \) be the first \( i > n \) for which \( S_i < k/2\sqrt{n} \). By condition (25) and the fact that \( S_{(\sigma-1)\wedge i} > k/2\sqrt{n} > c/n \), the sequence \( S_{\sigma \wedge i} \) is a submartingale. Decompose this into a mean-zero martingale and an increasing process. Summing equation (27) shows the variance of the martingale to be bounded in \( L^2 \) by \( b_2/n \). Then by using the one-sided Tschebysheff estimate \( \mathbf{P}(f - \mathbf{E}f < -s) \leq \text{Var}(f)/(\text{Var}(f) + s^2) \), the probability that the martingale ever reaches the interval \( [-\infty, -k/2\sqrt{n}] \) is at most \( 4b_2/(4b_2 + k^2) \). The martingale is a lower bound for the submartingale so the claim is proved.

(C) If \( S_n \) converges to 0 with non-zero probability, then there is an \( n \) which can be chosen arbitrarily large and an event \( A \in \mathcal{F}_n \) for which \( \mathbf{P}(S_n \to 0 \mid A) \) is arbitrarily close to 1. When it is greater than \( 1 - a/2 \), this contradicts (A) and (B).
Now assume (23) instead of (24) - (27). For any \( N_0 \), let \( \sigma \) be the first \( n \geq N_0 \) for which \( V(n) \) exits \( \mathcal{N} \) or one of the conditions (24) - (27) is violated; \( \sigma \) is a stopping time since the conditions are \( \mathcal{F}_n \)-measurable. Let \( \{ X_n^*, S_n^* : n \geq N_0 \} \) be any process that always satisfies (24) - (27) and is coupled to the process \( \{ X_n, S_n : n > N_0 \} \) so that the two processes are equal for \( n \leq \sigma \). Since \( S_n^* \) cannot converge to \( p \), \( S_n \rightarrow p \) implies \( \sigma < \infty \). For \( \epsilon > 0 \) let \( N_0 \) be chosen as in (23). Then with probability at least \( 1 - \epsilon \) either \( S_n \) does not converge to \( p \) or \( V(n) \) exits \( \mathcal{N} \) after time \( N_0 \) is at most \( \epsilon \). Since \( \epsilon \) is arbitrary, it follows that \( \mathbb{P}(V(n) \rightarrow p) = 0. \)

The last step in the proof of Theorem 1.2 is to establish Proposition 5.4. For any \( \epsilon > 0 \) and \( \gamma < 1 \), condition (26) may be satisfied by choosing \( N_0 \) at least as large as the \( N_0 \) in Proposition 5.2 (iii) (using the fact that \( \eta \) is Lipschitz). Also, (27) follows directly from Proposition 5.2 (i) for any \( \epsilon \). To prove (25), let \( \epsilon > 0 \) and use the bounds on \( \tau_n \) from Proposition 5.2 (ii) to get

\[
\mathbb{E}(X_{n+1}) = \mathbb{E}(S_{n+1}) - S_n \\
= \mathbb{E}(\eta(U(n) + |U(N + 1) - U(n)|)) - S_n \\
\geq \mathbb{E}(\eta(U(n)) + D_{U(n)}\eta[U(n + 1) - U(n)] + O(U(n + 1) - U(n))^2) - S_n \\
\quad \text{by Lemma 5.1} \\
= D_{U(n)}\eta\mathbb{E}[U(n + 1) - U(n)] + \mathbb{E}(O(U(n + 1) - U(n)^2)) \\
= D_{U(n)}\eta\mathbb{E}\left[ \frac{L(U(n))}{\tau_n}(\pi - I)U(n) + O(\tau_n^{-2}) \right] + \mathbb{E}(O(U(n) - U(n)^2)) \\
\quad \text{by Proposition 5.3} \\
= \frac{L(U(n))}{\tau_n}D_{U(n)}\eta((\pi - I)(U(n))) + O(\tau_n^{-2}) \\
\quad \text{since } U(n + 1) - U(n) \text{ is of order } \tau_n^{-1} \text{ and } \eta \text{ is Lipschitz} \\
\geq k_1 \frac{L(U(n))}{\tau_n}\eta(U(n)) + O(\tau_n^{-2}) \\
\geq \frac{c_1 S_n}{n} - \frac{c_2}{n^2} \\
\geq \frac{c_1 S_n}{n} - \frac{c_2}{n^2} \quad (28)
\]
for some $c_1, c_2 > 0$ by Proposition 5.2 (ii) with probability $1 - \epsilon$.

Thus there is a constant $c = c_2/c_1$ such that for $S_n > c/n$ the first term of (28) dominates. Hence (25) is true with probability at least $1 - \epsilon$.

Finally, to show (24), note that it suffices to show that $\mathbb{E}(X_{n+1}^2) \geq c_3/n^2$ for some $c_3$, assuming $\tau_n \leq \tau$. For, in the case that $S_n > c_2/c_1 n$, (25) holds, implying (24), while if $S_n \leq c_2/c_1$, (28) is at least $-2c_2 n^{-2}$ and the second term on the left hand side of (24) is at least $-4c_3^2/c_1 n^{-3}$ and for large enough $n$ this is dwarfed by the $\mathbb{E}(X_{n+1}^2)$ term.

Now a moment’s thought shows that $\mathbb{E}(X_{n+1}^2)$ must be at least order $n^{-2}$: from the nonvanishing second partials of $\eta^2$, it follows that there is a unit vector $w \in W$ such that $|\eta(v + rw) - \eta(v)| > Cr$ for some positive $C$ uniformly in $v$ in a neighborhood of $p$. There exists a positive multiple of $w$ and a fixed sequence of sites $\{2, \ldots, d\}$, such that if these are the sites visited between times $\tau_n$ and $\tau_{n+1}$, then $U(n + 1) - U(n)$ will be arbitrarily close to this multiple of $w$. This sequence of visits happens with positive probability; so (24) holds, establishing Proposition 5.4 and Theorem 1.2.

6 Examples and further questions

Example 1: Suppose $R_{ij} = 1 - \delta_{ij}$. The critical set $C$ contains just the centroids of the faces, and the degeneracy set $C_0$ is empty, so by Corollary 2.3, $V(n)$ converges almost surely to some point of $C$. It is easy to see that the centroids of all proper faces are linear nonmaxima. For example, if $p = (1/3, 1/3, 1/3, 0, 0, \ldots)$ then $N_i = 2/3$ for $i \leq 3$ and $1$ for $i > 3$. Thus Theorem 1.3 implies $V(n) \rightarrow (\frac{1}{d}, \ldots, \frac{1}{d})$ almost surely.
Example 2: Here is an example where \( \lim_{n \to \infty} \) is not deterministic. Suppose

\[
R = \begin{bmatrix}
3 & 1 & 1 \\
1 & 2 & 4 \\
1 & 4 & 2
\end{bmatrix}.
\]

All the minors of \( R \) are invertible and off-diagonal elements nonzero, so Corollary 2.3 applies and \( V(n) \) converges almost surely to a point of \( C \). The interior point \( (1, 1, 1)R^{-1} = (1/2, 1/4, 1/4) \) is unstable because \( R \) has two positive eigenvalues, so the probability of convergence there is zero. The critical points in the middle of two of the edges, \( (1/3, 2/3, 0) \) and \( (1/3, 0, 2/3) \) are linear nonmaxima as are the vertices \( (0, 1, 0) \) and \( (1, 0, 1) \), so the probability of convergence to each of these points is zero by Theorem 1.3. On the other hand, \( (1, 0, 0) \) is a local maximum for \( H \) as is \( (0, 1/2, 1/2) \), so by Theorem 1.4, it follows that \( P(V(n) \to (1, 0, 0)) = 1 - P(V(n) \to (0, 1/2, 1/2)) = a \) for some \( 0 < a < 1 \).

Example 3: Let \( G \) be a finite abelian group and let \( T \) be a set of generators for \( G \) closed under inverse. Let \( R \) be the incidence matrix for the Cayley graph of \( (G, T) \). By symmetry, the point \( p = (1/|G|, \ldots, 1/|G|) \) is in \( C \). The eigenvalues of \( R \) are just \( \lambda(\chi) \overset{\text{def}}{=} \sum_{g \in T} \chi(g) \), as \( \chi \) ranges over the characters of \( G \). If these are all nonzero, then \( p \) is the unique critical point in the interior of \( \triangle \). In this case, \( P(V(n) \to p) \) is zero or not according to whether \( \lambda(\chi) > 0 \) for any nontrivial character \( \chi \). In fact it is easy to verify that \( P(V(n) \to p) \) is always zero or one when the principal minors of \( R \) are invertible, by checking that the negativity of \( \lambda(\chi) \) for all nontrivial \( \chi \) implies that each other critical point is a linear nonmaximum.

There are many natural unanswered questions about the behavior of \( V(n) \). One could of course ask for rates of convergence, central limit behavior, etc., but I think it is more important both from a mathematical and a modeling point of view to try to extend the results already obtained so as to cover all matrices \( R \). For example, when \( R \) is a matrix of all ones, every point of \( \triangle \) is critical so Theorem 1.1 says nothing, while comparison to
a Polya urn model shows that \( V(n) \) converges almost surely to a random point of \( \triangle \) with an absolutely continuous distribution. In general, when \( C \) has components larger than a point, one expects the motion of \( V \) inside a component to be martingale-like and hence still converge to a single point, this time with a nonatomic distribution. Also, while the symmetry assumption on \( R \) is vital to the proofs (since it allows \( \pi(v) \) to be explicitly calculated) I do not believe that it is actually necessary for the results.

**Conjecture 1** \( \lim_{n \to \infty} V(n) \) exists almost surely without any nondegeneracy assumptions on \( R \).

**Conjecture 2** Theorem 1.1 holds whether or not \( R \) is symmetric. Also, when \( R \) is not symmetric, there is a function \( H \) such that the first part of Lemma 2.1 holds and Theorem 1.2 holds.

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