Genomic signatures of geographic isolation and natural selection in coral reef fishes

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Abstract

The drivers of speciation remain among the most controversial topics in evolutionary biology. Initially, Darwin emphasized natural selection as a primary mechanism of speciation, but the architects of the modern synthesis largely abandoned that view in favour of divergence by geographic isolation. The balance between selection and isolation is still at the forefront of the evolutionary debate, especially for the world’s tropical oceans where biodiversity is high, but isolating barriers are few. Here, we identify the drivers of speciation in Pacific reef fishes of the genus Acanthurus by comparative genome scans of two peripheral populations that split from a large Central-West Pacific lineage at roughly the same time. Mitochondrial sequences indicate that populations in the Hawaiian Archipelago and the Marquesas Islands became isolated approximately 0.5 Ma. The Hawaiian lineage is morphologically indistinguishable from the widespread Pacific form, but the Marquesan form is recognized as a distinct species that occupies an unusual tropical ecosystem characterized by upwelling, turbidity, temperature fluctuations, algal blooms and little coral cover. An analysis of 3737 SNPs reveals a strong signal of selection at the Marquesas, with 59 loci under disruptive selection including an opsin Rh2 locus. While both the Hawaiian and Marquesan populations indicate signals of drift, the former shows a weak signal of selection that is comparable with populations in the Central-West Pacific. This contrast between closely related lineages reveals one population diverging due primarily to geographic isolation and genetic drift, and the other achieving taxonomic species status under the influence of selection.

Keywords: adaptation, genetic drift, natural selection, outlier loci, RADSeq, speciation

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Introduction

Speciation, or the evolutionary process by which new species arise, continues to be one of the most controversial topics in evolutionary biology. Views on how species are formed have changed dramatically in the last 150 years. Darwin’s idea that new species form by disruptive selection was heavily criticized by the architects of the modern synthesis and almost abandoned in view of species originating by geographic isolation (Darwin 1859; Coyne & Orr 2004). However, with the support of numerous examples, ecological speciation is now widely accepted (Schluter 2009; Nosil 2012; Bowen et al. 2013) but the relative importance of geographic isolation, and the genetic changes that lead to speciation, is still hotly debated topic in evolutionary biology (Nosil et al. 2009; Via 2009; Supple et al. 2014). With recent access to genomic information, the traditional view of a cohesive, gradually evolving genome is giving way to a
view of permeable genomes, wherein islands of genomic differentiation are linked to loci under disruptive selection (Wu 2001; Via 2012). This pattern is usually associated with sympatric speciation (Michel et al. 2010); however, ecological speciation is not tied to a specific geographical context (Rundle & Nosil 2005). Furthermore, natural selection can drive populations to reproductive isolation with or without gene flow (Schluter 2001; Nosil 2012).

Marine organisms are a particular challenge to models of speciation that require isolation for extended periods of time as many of these taxa have large population sizes, vast geographic ranges and long pelagic larval phases that extend for weeks to months, potentially connecting distant habitats (Bowen et al. 2013). Furthermore, there is a relative paucity of strong barriers to dispersal in the oceans, with recognized barriers often being porous in nature and sporadically traversed when conditions are favourable (Briggs & Bowen 2013; Gaither & Rocha 2013). These observations, combined with the high biodiversity on coral reefs, prompt the conclusion that natural selection, perhaps reinforced by assortative mating (e.g. Puebla et al. 2014), must play an important role in speciation in the sea.

A prerequisite for assortative mating is mate recognition, and among the often brilliantly coloured fishes found on coral reefs, visual cues are thought to play a key role (Helfman et al. 2009). While colour patterns are expected to diverge slowly under genetic drift, they can evolve rapidly when this character is involved in mate choice (Endler & Houde 1995; McMillan et al. 1999; Puebla et al. 2007, 2014). In the Caribbean reef fishes known as hamlets (genus Hypoplectrus), distinct colour morphs that occur in isolating populations (Salzburger et al. 2006; Blais et al. 2009; Selz et al. 2014). In both cases, shifts in colour patterns driven by natural selection outpace genetic divergence at neutral loci; however, there are also abundant counter-examples (McCartney et al. 2003; Rocha 2004; Messmer et al. 2005; Craig et al. 2006; Drew & Barber 2009; Leray et al. 2010; DiBattista et al. 2012).

With high species richness and recent evolutionary radiations (Choat et al. 2012; Sorenson et al. 2013; Gaither et al. 2014), tropical oceans are an excellent setting for the study of ecological speciation, yet there are only a few case studies (Munday et al. 2004; Rocha et al. 2005; Johannesson 2009; Bird et al. 2011; Prada & Hellberg 2013). Evidence for selection has been inferred from shifts in habitat use (Munday et al. 2004; Bird et al. 2011), differences in survivorship (Prada & Hellberg 2013) or divergences at neutral loci that correspond with ecological rather than geographic partitions (Rocha 2004). Few studies have yet provided the genomic evidence to support these inferences, but initial results are promising. In hamlets, genome scans indicate that Hox gene evolution may be involved in the radiation of distinct morphotypes (Puebla et al. 2014). In freshwater fishes, compelling cases of divergence driven by natural selection are found in three-spine sticklebacks (Gasterosteus aculeatus), where genomic islands of divergence correspond to colonization of novel habitats (Hohenlohe et al. 2010) and in lake whitefish (Coregonus clupeformis), where divergence coincides with niche partitioning (Renaut et al. 2012). Few studies, however, have explored the contrast between the genomic signatures of differentiation caused by geographic isolation and disruptive selection in marine organisms.

To resolve the genomic signatures of isolation and selection, we analysed two members of the Orangeband Surgeonfish species complex: Acanthurus olivaceus Bloch and Schneider 1801 and A. reversus Randall and Earl 1999. The widely distributed A. olivaceus occurs across the central and western Pacific and extends to the Hawaiian Archipelago, while A. reversus, recently described as a separate species, is endemic to the Marquesas (Fig. 1, Randall & Earle 1999). Only slight differences in fin ray counts distinguish A. reversus from A. olivaceus (Table S1, Supporting information), but they can easily be differentiated based on colour patterns (Randall 2005; Fig. 1a). These species, like most surgeonfishes, inhabit shallow waters around coral reefs (Robertson 1983). It is estimated that A. olivaceus can live for at least 35 years (Choat & Axe 1996), yet as in most coral reef organisms dispersal is limited to a pelagic larval stage (Leis 1991; Leis & McCormick 2002). The pelagic larval duration for our study species has not been determined but estimates for other Acanthurids range from 40 to 70 days (Doherty et al. 1995; Fisher et al. 2005).

Initially a targeted locus approach, including mtDNA cytochrome b (Cytb), and nuclear introns S7 and Gpd2 (data reported here), revealed that the isolated Hawaiian population of A. olivaceus began diverging from the Central-West Pacific populations at roughly the same time as the Marquesan A. reversus: a surprising result given the taxonomic status and unique coloration of A. reversus (Fig. 1a). In light of these findings, we analysed 3737 single nucleotide polymorphisms (SNPs) obtained through RADSeq (Restriction Site Associated DNA Sequencing) (Davey et al. 2011; Peterson et al. 2012). These genome-wide surveys were used to detect signals of selection and isolation in the Hawaiian and Marquesan lineages, to answer two questions: Is it possible to distinguish between isolation and selection in these recently diverged species? Has selection accelerated evolutionary divergence at the Marquesas?
Materials and methods

Specimen collection

A total of 302 specimens of *A. olivaceus* and 30 specimens of *A. reversus* were collected either by SCUBA divers using pole spears or obtained from local fish markets where the provenance of the fish was known (Table 1). Gill, fin and/or muscle tissues were preserved in either salt-saturated DMSO or 95% ethanol and stored at room temperature at the Hawai‘i Institute of Marine Biology.

Sanger sequence analyses

Laboratory protocols for Sanger sequencing are provided in Supplemental information. Summary statistics for the mtDNA, including haplotype diversity (*h*) and nucleotide diversity (*π*), were estimated using algorithms from Nei (1987) as implemented in the software package Arlequin 3.5 (Excoffier et al. 2005). The AIC implemented in jMODELTEST 0.1.1 (Posada 2008) indicated the TPM3uf+G as the best fit model of DNA sequence evolution with a gamma value of 0.32. Median-joining networks were constructed using the program NETWORK 4.5 with default settings (Bandelt et al. 1999). To evaluate the phylogenetic relationship between the study species, we used the maximum-likelihood (ML) method and default settings of the program MEGA 5.0 (Tamura et al. 2011). Cytb sequences for surgeonfishes of the genus *Acanthurus* (Acanthuridae) were obtained from GenBank (Table S2, Supporting information) including the closely related *A. tennentii* (Sorenson et al. 2013), and trees were rooted using sequence data from the Spotted Surgeonfish, *Ctenochaetus strigosus* and the Elongate Unicornfish, *Naso lopezi*. Bootstrap support values were calculated using default settings with 2000 replicates. The ML tree topology was confirmed by Bayesian Markov Chain Monte Carlo (MCMC) analysis using MrBayes 3.1.1 (Ronquist & Huelsenbeck 2003). The Bayesian analysis was run using the GTR model with gamma-distributed rate variation across sites for two million iterations sampling every 100th generation with an initial burn-in of 5000 samples. We calculated the divergence between mitochondrial lineages (*d*) in Arlequin.

To test for genetic structure within and between species, an analysis of molecular variance (AMOVA) was performed in Arlequin using 20 000 permutations. Because the TPM3uf+G model of sequence evolution is not implemented in Arlequin, we used the most similar model available (TrN; Tamura & Nei 1993) with a gamma value of 0.32. An analogue of Wright’s *F* bar, *Φ* bar, which incorporates a model of sequence evolution and distance, was calculated for the entire data set and for pairwise comparisons among all locations. We maintained *α* = 0.05 among all pairwise tests by controlling for the false discovery rate (Narum 2006).

To estimate the time to coalescence based on the mtDNA data set, we used the Bayesian MCMC approach implemented in BEAST 1.7.4 (Drummond & Rambaut 2007). We conducted our analysis with a relaxed lognormal clock and uncorrelated substitution rates among branches. We used default priors under the HKY + G model of mutation and ran simulations for 10 million generations with sampling every 1000
Table 1 Molecular diversity indices for populations of Acanthurus olivaceus and A. reversus

| Sample location  | Cyth     |       |       |       |       |       |       |       |       |       |       |       |       |       |
|------------------|----------|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|-------|
|                  | N   | N_e | h   | π   |       |       |       |       |       |       |       |       |       |       |
| Acanthurus olivaceus |       |       |       |       |       |       |       |       |       |       |       |       |       |       |
| Hawai’i (HI)     | 35  | 3   | 0.51 ± 0.05 | 0.003 ± 0.002 | 32 | 6 | 0.69 | 0.62 | 0.958 | 35 | 6 | 0.57 | 0.65 | 0.358 |
| Palmyra (PA)     | 26  | 11  | 0.75 ± 0.09 | 0.004 ± 0.003 | 30 | 7 | 0.73 | 0.71 | 0.735 | 29 | 4 | 0.69 | 0.68 | 0.314 |
| Kirimiti (KI)    | 32  | 16  | 0.86 ± 0.05 | 0.005 ± 0.003 | 32 | 10 | 0.74 | 0.70 | 0.906 | 32 | 5 | 0.56 | 0.57 | 0.876 |
| Moorea (MO)      | 25  | 10  | 0.88 ± 0.04 | 0.005 ± 0.003 | 29 | 4 | 0.62 | 0.62 | 0.874 | 24 | 3 | 0.63 | 0.64 | 0.384 |
| Cook Islands (CO)| 27  | 6   | 0.80 ± 0.04 | 0.003 ± 0.002 | 29 | 5 | 0.57 | 0.57 | 0.918 | 27 | 3 | 0.48 | 0.65 | 0.065 |
| Fiji (FI)        | 7   | 5   | 0.86 ± 0.14 | 0.004 ± 0.003 | 8  | 7  | 0.75 | 0.80 | 0.561 | 7  | 3  | 0.86 | 0.63 | 0.663 |
| Marshall Islands (MI) | 25 | 8   | 0.75 ± 0.06 | 0.003 ± 0.002 | 26 | 12 | 0.62 | 0.80 | 0.016 | 26 | 3 | 0.77 | 0.66 | 0.598 |
| Palau (PL)       | 29  | 13  | 0.84 ± 0.06 | 0.004 ± 0.003 | 21 | 8  | 0.67 | 0.69 | 0.098 | 27 | 6 | 0.59 | 0.71 | 0.583 |
| Philippines (PP) | 18  | 8   | 0.80 ± 0.09 | 0.006 ± 0.003 |       |       |       |       |       |       |       |       |       |       |
| Bali (BA)        | 13  | 7   | 0.79 ± 0.11 | 0.005 ± 0.003 | 10 | 5  | 0.70 | 0.70 | 0.683 | 8  | 3 | 0.75 | 0.70 | 0.061 |
| Christmas Island (CI) | 37 | 13  | 0.68 ± 0.09 | 0.004 ± 0.002 | 32 | 8  | 0.71 | 0.72 | 0.117 | 36 | 6 | 0.69 | 0.67 | 0.499 |
| Cocos-Keeling (CK)| 15  | 7   | 0.86 ± 0.06 | 0.005 ± 0.003 | 17 | 3  | 0.44 | 0.45 | 1.000 | 16 | 3 | 0.69 | 0.69 | 0.628 |
| Acanthurus reversus |       |       |       |       |       |       |       |       |       |       |       |       |       |       |
| Marquesas (MQ)   | 30  | 13  | 0.80 ± 0.07 | 0.003 ± 0.002 | 26 | 6  | 0.73 | 0.68 | 0.610 | 27 | 3 | 0.56 | 0.58 | 0.183 |

Sample location, number of individuals (N), number of haplotypes (N_h), haplotype diversity (h) and nucleotide diversity (π) are listed for the cytochrome b (Cyth) data set. Number of alleles (N_a), observed heterozygosity (H_o) and expected heterozygosity (H_e) are listed for the S7 ribosomal protein gene (S7) and the glyceraldehyde-3-phosphate dehydrogenase (Gpd2) nuclear introns. P-values are the result of exact tests for Hardy–Weinberg equilibrium using a Markov chain with 100 000 steps in ARLEQUIN. We maintained α = 0.05 among all pairwise tests by controlling for the false discovery (Narum 2006). The corrected α = 0.015.

generations. A sequence divergence estimate of 2% per My between lineages was used to estimate coalescence times (Bowen et al. 2001; Reece et al. 2010). Five independent runs were computed to ensure convergence, and log files were combined using the program TRACER 1.5 (http://tree.bio.ed.ac.uk/software/tracer/).

Allelic states of nuclear intron sequences were estimated using the Bayesian program PHASE 2.1 (Stephens & Donnelly 2003) as implemented in the software DNASP 5.0 (Librado & Rozas 2009; see Supplemental information). Using the results from PHASE, observed heterozygosity (H_o) and expected heterozygosity (H_e) were calculated for each nuclear intron and an exact test of Hardy–Weinberg equilibrium (HWE) using 100 000 steps in a Markov chain was performed using ARLEQUIN. Linkage disequilibrium between the two loci was assessed using the likelihood ratio test with 20 000 permutations in ARLEQUIN. Genotypes for each individual at the S7 and Gpd2 introns were compiled and used to calculate FST for the multilocus data set and for pairwise comparisons between locations in ARLEQUIN. The false discovery rate among multiple comparisons was controlled as described above. Median-joining networks for alleles were constructed for each locus using NETWORK.

RADSeq data

The details of library preparation are provided in Supplemental information. Libraries were sequenced at the UCLA Neuroscience Genomics Core facility on an Illumina HiSeq 2000 (100 bp single end reads). Sequencing runs resulted in 135 million reads passing initial quality control at the sequencing facility. Raw sequences were de-multiplexed, trimmed to a common length, and reads with Phred scores below 33 were discarded, resulting in 132 million reads with the number varying between 4.3 thousand and 4.6 million reads per individual. Nine individuals with <60 thousand reads were discarded, leaving 61 individuals across six populations (A. olivaceus: Hawai’i, N = 14; Kirimiti, N = 12; Marshall Islands, N = 7; Palau, N = 8; Bali, N = 11; A. reversus: Marquesas, N = 9). Loci were assembled using the STACKS denovo_map.pl pipeline and its component programs v. 1.09 (Catchen et al. 2011). Loci were generated by the merging of three or more similar ‘stacks’. Stacks in which the number of reads was more than two standard deviations above the mean were assumed to be repetitive elements and removed from analysis. Loci were identified by aligning the sequence reads from each individual to the assembled contigs resulting in 844 220 stacks. Using the ‘populations’ component in STACKS, we further filtered the data set retaining only those loci with ≥8× coverage (‘m’ command) and that aligned in ≥70% of individuals (‘r’ command) in each population (‘p’ command). The resulting data set consisted of 3737 loci with a data matrix that was 88.5% complete. From this data set, GENEPOP 4.2 (Rousset 2008) and STRUCTURE 2.3.4 (Pritchard et al. 2000) input files.
were produced using the ‘populations’ component in STACKS and implementing the ‘write_single_snp’ option. The resulting files were used for population level analyses or converted into other formats using the program PGDSPIDER 2.0 (Lischer & Excoffier 2012).

We tested for outlier loci across all populations using the FDIST approach (Beaumont & Nichols 1996) as implemented in ARLEQUIN. This method estimates the expected distribution of Wright’s inbreeding coefficient $F_{ST}$ vs. expected heterozygosity under an island model of migration. Outlier loci are those that show either significantly higher (divergent selection) or lower (balancing selection) $F_{ST}$ compared to simulated neutral expectations. We ran 20 000 simulations with 100 demes per group (number of groups equals the number of populations ran plus 1) and minimum and maximum expected heterozygosities of 0 and 0.8, respectively. Based on the results of this analysis, we conservatively classified each locus into one of three selection categories: neutral (between the 95% quantiles), balancing (below the 1% quantile) or divergent (above the 99% quantile) selection. Under this classification scheme, 9.7% (362) of the loci remained unclassified. We also detected outlier loci using BAYESCAN 2.1. This program uses a logistic regression model that implements a locus expectation that had a sequence match of ≥50% and an $E$-value of ≤ e-10.

To test for population structure, an analysis of molecular variance (AMOVA) was performed in ARLEQUIN using 99 999 permutations. $F_{ST}$ was calculated for the overall data set and for intraspecific pairwise comparisons among locations. Genetic partitioning was assessed using STRUCTURE with an initial burn-in of 50 000 replicates, followed by 500 000 replicates for each population ($K$). The analyses were run for $K = 1$–6, each replicated 10 times, without population priors. We determined the most likely number of genetic clusters ($K$) based on the results of STRUCTURE HARVESTER 0.6.93 (Earl & von Holdt 2012). STRUCTURE results were analysed using the program CLUMPP 1.1.2 (Jakobsson & Rosenberg 2007), which minimizes the variance across all iterations, and the final output was visualized using the software DISTRUCT (Rosenberg 2004). Results from these preliminary runs never returned a $K > 3$; therefore, we preformed 10 final runs of 500 000 replicates burn-in and 1 000 000 replicates for each of $K = 1$–4.

Finally, we conducted principle coordinate analyses (PCoA) using GENALEX 6.5 (Peakall & Smouse 2012). This technique plots major patterns within multivariate data sets. A genetic distance matrix was calculated for each data set using the codominant (genotypic) data option. All four methods of calculation available in GENALEX (covariance-standardized, covariance-not standardized, distance-standardized and distance-not standardized) were tested and produced similar patterns. Here, we present the results of the covariance matrix without data standardization.

To resolve the function of loci under divergent selection, we searched these loci against the publicly available genomes using the NCBI Blast plug-in of GENEIOUS PRO (Biomatters). We used the programs MegaBlast and Blastn and, as expected, obtained more hits with the former. Here, we report the Blastn top hit for each fragment that had a sequence match of ≥50% and an $E$-value of ≤ e-10.

Results

Divergence of peripheral populations

We resolved 509 bp of Cytb in 319 individuals yielding 59 haplotypes (Fig. 2a, Table 1). Nucleotide diversity among populations was low ($\pi = 0.003$–0.006), while the corresponding haplotype diversity was high ($h = 0.51–0.88$; Table 1). Phylogenetic analysis of Cytb data for surgeonfishes (Acanthuridae; Table S2, Supporting information) supports the sister relationship between A. olivaceus and A. reversus (Fig. 1c and Fig. S1, Supporting information), with no diagnostic (fixed) differences between species. A median-joining network revealed a tight cluster of haplotypes, many of which are shared between putative species (Fig. 2a). Overall population structure was $\Phi_{ST} = 0.103$ ($P < 0.001$). Most of this structure was due to distinct populations at Hawaii’i and the Marquesas (Table 2). Surprisingly, we found higher genetic structure distinguishing Hawaii’i from conspecifics in the Central-West Pacific (pairwise $\Phi_{ST} = 0.163–0.392$) compared with population structure between A. reversus and central Pacific A. olivaceus (pairwise $\Phi_{ST} = 0.115–0.262$; Table 2). Apart from Hawaii’i and the Marquesas, we detected only a single significant comparison across the remainder of the range (Christmas Island vs. Cook Islands; $\Phi_{ST} = 0.077$, $P = 0.004$).

We also resolved 171 bp of the S7 intron in 292 specimens (15 variable sites resulting in 22 alleles, Fig. 2b) and 76 bp of the Gpd2 intron in 294 specimens (five variable sites resulting in seven alleles; Fig. 2c). The number of individuals (N), number of alleles (Na), observed heterozygosity ($H_0$), expected heterozygosity ($H_E$) and the corresponding $P$-value for the exact test for HWE for each locus are provided in Table 1. Despite multiple attempts, we were unable to amplify the samples from the Philip-
Fig. 2 Median-joining networks for (a) mtDNA cytochrome b, (b) the first intron of the S7 ribosomal protein and (c) intron 2 of the Glyceraldehyde-3-phosphate dehydrogenase (Gpd2) gene for *Acanthurus olivaceus* and *A. reversus*. Each circle represents one mitochondrial haplotype or nuclear intron allele with the area of each circle proportional to haplotype or allele frequency; small black circles represent missing haplotypes; to simplify the figure, locations with >4 singletons branching from a single haplotype were tallied and recorded in parenthesis; colours represent collection location (see key).

### Table 2 Pairwise ΦST values for *Acanthurus olivaceus* and *A. reversus* populations

| Location         | 1   | 2   | 3   | 4   | 5   | 6   | 7   | 8   | 9   | 10  | 11  | 12  | 13  |
|------------------|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|
| *Acanthurus*     |     |     |     |     |     |     |     |     |     |     |     |     |     |
| olivaceus        |     |     |     |     |     |     |     |     |     |     |     |     |     |
| 1. Hawai‘i       | 0.000 | 0.001 | 0.000 | 0.000 | 0.002 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 | 0.000 |
| 2. Palmyra       | 0.327 | 0.216 | 0.023 | 0.030 | 0.828 | 0.468 | 0.538 | 0.671 | 0.724 | 0.449 | 0.499 | 0.000 | 0.000 |
| 3. Kiritimati     | 0.163 | 0.009 | 0.262 | 0.024 | 0.342 | 0.674 | 0.974 | 0.613 | 0.438 | 0.286 | 0.393 | 0.000 | 0.000 |
| 4. Moorea        | 0.234 | 0.047 | 0.007 | 0.550 | 0.242 | 0.153 | 0.148 | 0.100 | 0.088 | 0.010 | 0.175 | 0.000 | 0.000 |
| 5. Cook Isl.     | 0.376 | 0.047 | 0.046 | −0.008 | 0.225 | 0.081 | 0.031 | 0.032 | 0.077 | 0.004 | 0.136 | 0.000 | 0.000 |
| 6. Fiji          | 0.392 | −0.039 | 0.006 | 0.024 | 0.027 | 0.330 | 0.395 | 0.734 | 0.770 | 0.527 | 0.681 | 0.000 | 0.000 |
| 7. Marshall Isl. | 0.275 | −0.004 | −0.012 | 0.018 | 0.033 | 0.005 | 0.858 | 0.454 | 0.529 | 0.299 | 0.524 | 0.000 | 0.000 |
| 8. Palau         | 0.205 | −0.006 | −0.021 | 0.017 | 0.046 | 0.001 | −0.019 | 0.563 | 0.594 | 0.640 | 0.337 | 0.000 | 0.000 |
| 9. Philippines   | 0.268 | −0.015 | −0.011 | 0.031 | 0.056 | −0.033 | −0.003 | −0.009 | 0.807 | 0.292 | 0.602 | 0.000 | 0.000 |
| 10. Bali         | 0.309 | −0.017 | −0.004 | 0.037 | 0.048 | −0.038 | −0.010 | −0.013 | −0.029 | 0.366 | 0.599 | 0.000 | 0.000 |
| 11. Christmas Isl.| 0.271 | −0.003 | 0.004 | 0.059 | 0.077 | −0.014 | 0.004 | 0.010 | 0.006 | 0.000 | 0.113 | 0.000 | 0.000 |
| 12. Cocos-Keeling| 0.294 | −0.006 | −0.001 | 0.020 | 0.028 | −0.027 | −0.009 | 0.003 | −0.011 | −0.013 | 0.030 | 0.000 | 0.000 |
| *A. reversus*    |     |     |     |     |     |     |     |     |     |     |     |     |     |
| 13. Marquesas    | 0.397 | 0.208 | 0.161 | 0.115 | 0.158 | 0.244 | 0.166 | 0.185 | 0.192 | 0.197 | 0.262 | 0.133 | 0.000 |

Cytochrome b data are below diagonal, and corresponding P-values are above the diagonal. Significant comparisons are underlined.

We maintained $\alpha = 0.05$ among all pairwise tests by controlling for the false discovery (Narum 2006). The corrected $\alpha = 0.01$. 

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pines at the nuclear introns, most likely due to template quality. Following correction for multiple comparisons, no samples significantly deviated from HWE expectations and no linkage disequilibrium was detected between the two loci \((P > 0.05)\). Similar to \textit{Cytb}, the median-joining networks for the introns revealed tight clusters of closely related alleles that are shared between species (Fig. 2). The two most common alleles at either

\begin{table}
\centering
\begin{tabular}{lccccccc}
\hline
 & 1 & 2 & 3 & 4 & 5 & 6 \\
\hline
\textit{Acanthus olivaceus} & & & & & & \\
1. Hawai‘i & & & & & & \\
2. Kiritimati & 0.023 & <0.001 & <0.001 & <0.001 & <0.001 & <0.001 \\
3. Palau & 0.020 & -0.002 & 0.065 & 0.486 & 0.011 & <0.001 \\
4. Marshall Isl. & 0.011 & -0.009 & -0.022 & 0.924 & 0.005 & <0.001 \\
5. Bali & 0.027 & 0.001 & 0.001 & 0.001 & 0.001 & 0.009 \\
& & & & & & \\
\textit{A. reversus} & & & & & & \\
6. Marquesas & 0.114 & 0.102 & 0.093 & 0.094 & 0.108 & \\
\hline
\end{tabular}
\caption{Pairwise \(F_{ST}\) values for \textit{Acanthus olivaceus} and \textit{A. reversus} populations based on 3737 RADSeq loci (SNPs)}
\end{table}

Corresponding \(P\)-values are above the diagonal. We maintained \(\alpha = 0.05\) among all pairwise tests by controlling for the false discovery (Narum 2006). The corrected \(\alpha = 0.014\). Significant comparisons are underlined.

\begin{table}
\centering
\begin{tabular}{lcccccc}
\hline
 & 1 & 2 & 3 & 4 & 5 & 6 \\
\hline
\textit{Acanthus olivaceus} & & & & & & \\
1. Hawai‘i & & & & & & \\
2. Kiritimati & 0.029 & 0.078 & 0.103 & 0.107 & 0.196 & 0.548 \\
3. Palau & 0.032 & 0.000 & 0.035 & 0.046 & 0.168 & 0.552 \\
4. Marshall Isl. & 0.027 & 0.005 & -0.003 & 0.004 & 0.006 & 0.526 \\
5. Bali & 0.033 & -0.003 & 0.004 & 0.006 & 0.140 & 0.478 \\
& & & & & & \\
\textit{A. reversus} & & & & & & \\
6. Marquesas & 0.114 & 0.098 & 0.094 & 0.101 & 0.106 & \\
\hline
\end{tabular}
\caption{Pairwise \(F_{ST}\) values for RADSeq loci (SNPs) for \textit{Acanthus olivaceus} and \textit{A. reversus} populations based on 1525 neutral loci (below diagonal) and 59 loci under divergent selection (above diagonal)}
\end{table}

We maintained \(\alpha = 0.05\) among all pairwise tests by controlling for the false discovery (Narum 2006). The corrected \(\alpha = 0.014\). Significant comparisons are underlined.
locus were observed in both species and at every location, except Cocos-Keeling (S7 locus). We detected no population structure in the combined (multi-locus) nuclear intron data set (overall $F_{ST} = 0.002$, $P = 0.185$), and none of the pairwise comparisons were significant after correction for multiple comparisons (data not shown). When the introns were evaluated separately, overall population structure in the S7 data set was low but significant ($F_{ST} = 0.009$, $P = 0.038$) with only a single significant pairwise comparison (Hawai‘i vs. Marquesas, $F_{ST} = 0.048$, $P = 0.005$; Table S3, Supporting information). No structure was detected with Gpd2 intron ($F_{ST} = 0.009$, $P = 0.038$; Table S3, Supporting information).

**Coalescent times**

To estimate divergence time, we conducted coalescent analysis on the complete mtDNA data set, which resulted in an estimated time to most recent common ancestor (TMRCA) for *A. olivaceus* and *A. reversus* of 0.83 Ma (95% HPD = 0.45–1.36 Ma). When analysed separately, TMRCA for *A. olivaceus* was 0.68 Ma (95% HPD = 0.41–0.99 Ma) and 0.49 Ma for *A. reversus* (95% HPD = 0.28–0.77 Ma). When only the Hawaiian population of *A. olivaceus* was considered, TMRCA = 0.55 Ma (95% HPD = 0.34–0.79 Ma). Central-West Pacific populations of *A. olivaceus* were similar (TMRCA = 0.58 Ma, 95% HPD = 0.38–0.82 Ma). While caution is warranted because these dates were calculated based on a single mtDNA marker, they imply that *A. reversus* began diverging from the Central-West Pacific population (0.49 Ma) at about the same time as the Hawaiian population of *A. olivaceus* (0.55 Ma).

**Genome scans**

We amplified 3737 RADSeq loci in 61 individuals across six sample locations (Fig. 1b) producing a robust data matrix that was 88.5% complete. Using the ‘–write_single_snp’ command in STACKS (as described above), we analysed the first single nucleotide polymorphism (SNP) in each locus. Employing the FDIST approach across all populations (ARLEQUIN), we detected 1525 neutral loci, 1791 loci under balancing selection, and 59 loci under divergent selection (pairwise $F_{ST}$ = 0.035–0.552; Table 4, Fig. 3). Under this conservative classification scheme, 9.7% (362) of the loci remained unclassified. We detected 16 outlier loci using BAYESSCAN, which were a subset of the ones detected using the FDIST approach with the exception of a single locus (Table S4, Supporting information). When the Marquesas population was removed from the analysis, the number of loci under divergent selection detected using the FDIST approach dropped by 43% (to 34 loci) with lower $F_{ST}$ values (pairwise $F_{ST}$ = 0.14–0.39, data not shown). For all subsequent analyses, we used the 59 outlier loci detected using the FDIST approach.

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contrast to the mtDNA data set we detected low but significant genetic structure in the population of *A. olivaceus* at Bali (Table 3). When the loci under balancing selection were considered, we found no population structure (overall $F_{ST} = -0.031, P = 1.000$). In contrast, when the neutral loci were analysed, we found significant population structure ($F_{ST} = 0.043, P < 0.001$; Table 4). Pairwise $F_{ST}$ values were highest for the Marquesan population ($F_{ST} = 0.094–0.106$) and lower for Hawai’i ($F_{ST} = 0.027–0.033$). Bali demonstrated low but significant structure when compared to both Palau ($F_{ST} = 0.004, P < 0.001$) and the Marshall Islands ($F_{ST} = 0.006, P < 0.001$). For loci under divergent selection, overall $F_{ST}$ was high and significant ($F_{ST} = 0.313, P < 0.001$) and all pairwise $F_{ST}$ values were significant. Population structure was highest for the Marquesan population (pairwise $F_{ST} = 0.478–0.552$), while values from Hawai’i (pairwise $F_{ST} = 0.078–0.196$) and within the Central-West Pacific were much lower (pairwise $F_{ST} = 0.035–0.168$; Table 4).

Results from the STRUCTURE analyses indicate $K = 2$ (3737 loci) with a lack of clustering among populations that is supported by the PCoA (Fig. 4a). Analyses of the neutral loci (Fig. 4b, $K = 2$) and the loci under balancing selection (Fig. 4c, $K = 2$ and 3 had similar probabilities and lack of clustering; $K = 2$ is shown) resulted in similar patterns with no partitioning among populations. In contrast, when the 59 outlier loci were analysed, STRUCTURE resulted in $K = 2$ with *A. reversus* clearly distinguished from *A. olivaceus*, but with no corresponding distinction of the Hawaiian population. The distinction of the Marquesan population was supported by PCoA only when the outlier loci were considered (Fig. 4d). The results of our NCBI Blast searches mapped one of the outlier loci to the opsin Rh2 gene (retinal pigment Rh 2A alpha; locus 65 948; Table S4, Supporting information). There were two alleles detected at this locus that are in nearly equal frequency across the data set (average gene diversity $= 0.49 \pm 0.47$). However, allele frequencies at the Marquesas were nearly fixed for one variant (90%).

**Discussion**

While we have a theoretical understanding of the roles that natural selection and geographic isolation play in shaping biodiversity, empirical studies of these processes in the sea are rare. The Orangeband Surgeonfish species complex forms a monophyletic lineage (Fig. 1c) and provides an unusual opportunity to address the role of selection in speciation with one geographically isolated population in the Hawaiian Islands and a sister species (*A. reversus*) of similar age (0.55 and 0.49 Ma), endemic to the Marquesas Islands. Using a targeted locus approach (mtDNA and nuclear introns), we found no diagnostic differences between these two lineages (Fig. 2). However, we did observe a strong shift in mtDNA haplotype frequencies (Table 2) with no accompanying signal of divergence in nuclear introns (Table S3, Supporting information). While little population structure was detected among *A. olivaceus* populations across the Central-West Pacific, the remote Hawaiian Archipelago was distinct and demonstrated population genetic divergence in mtDNA (shifts in haplotype frequencies) similar to *A. reversus* in the Marquesas Islands (Fig. 2 and Table 2). In contrast, our genome-wide survey revealed low levels of population structure at Bali not detected in the Sanger data sets (Tables 3 and 4). Most interesting was the detection of signals of divergent selection that distinguish the Marquesas, a signal not observed in the Hawaiian population of *A. olivaceus*, but with dozens of loci under divergent selection distinguishing *A. reversus*. This pattern is in concordance with the expectations of geographical isolation of Hawai’i, versus divergent selection acting upon lineages in the Marquesas.

Our data set does not negate the role of selection in geographically isolated populations such as those in Hawai’i nor does it dismiss the role of drift in populations under strong selection. On the contrary, our data set indicates that drift at neutral loci is important in both populations (Table 4), but it also highlights the significance of selection in disparate environments such as those observed in the Marquesas, adding to the accumulating evidence that selection in contrasting environments can be an important driver of speciation in the sea.

**Isolation in Hawai’i and selection at the Marquesas**

As expected from traditional views of allopatric speciation, much of the high rates of endemism in the oceans are observed in geographically isolated locations, such as oceanic islands (Briggs & Bowen 2012). The Hawaiian Archipelago, which has the highest level of endemism in Oceania (25% of shore fishes; Randall 2007), is one of the most remote island groups on Earth. Hawai’i is isolated to the east by a 4000 km stretch of open ocean. While Johnston Atoll lies just 850 km to the southwest, this island is small (2.7 km²) and provides little habitat. The small islands and atolls of the Line Islands lie about 1600 km to the south of Hawai’i but their northernmost atolls, Kingman Reef and Palmyra, are swept by equatorial currents that flow in a predominantly east–west direction. As a result, Hawai’i is exceptionally isolated and colonization of this archipelago likely results from infrequent dispersal, yielding high endemism and genetically distinct
populations (Craig et al. 2010; Bird et al. 2011; Gaither et al. 2011). Similar to Hawaii, the Marquesas are also isolated to the east. However, just 475 km southwest of the Marquesas lies a network of islands that extends to Australia, with distances no greater than 800 km between shallow habitats (Schultz et al. 2008). Prevailing currents in the region flow in a westerly direction, but genetic studies indicate that dispersal in the opposing direction is frequent enough to preclude the development of population structure among some reef fishes (Gaither et al. 2010). Thus, the high rate of endemism at these islands is difficult to explain by geographic isolation alone. Instead, what distinguish the Marquesas are unique environmental conditions. Unlike most tropical Pacific islands, the Marquesas lack fringing reefs and lagoons. These islands are exposed to unusually variable sea temperatures for an equatorial location with temperatures ranging from 26 to 30 °C in the open ocean, while there are reports of temperatures as low 15–18 °C due to upwelling (Randall & Earle 2000). Additionally, there are frequent phytoplankton blooms and restricted coral development (Randall & Earle 2000; Martinez & Maamaatuaiahutapu 2004). The lack of lagoon habitat and reduced coral cover, combined with variable sea water temperatures and high nutrient input, is thought to drive the biogeographic distinction of the Marquesas, which harbours unusually low species richness and high endemism (12% of shore fishes; Randall 2005). Furthermore, several phylogeographic surveys of widely distributed species have shown significant population differentiation between the Marquesas and the Central Pacific (Planes & Fauvelot 2002; Gaither et al. 2010; Ludt et al. 2012; Szabó et al. 2014). Colonists that arrive here are likely to be under strong selection in these unique ecological conditions, which may lead to differentiation of lineages and, ultimately, to speciation.

Colour as a taxonomic trait

The delineation of species is fundamental to many aspects of evolutionary biology, systematics, biogeography and conservation. Until recently, most species were distinguished based on morphology; however, molecular technologies arising in the past few decades have revealed the shortcomings of using morphology as the sole criteria for taxonomy. An exceptional example is the bonefish Albula vulpes. Based on morphology, this species was considered a single globally distributed species until analysis of molecular data revealed deep genetic partitions in Hawaiian populations (Shaklee & Tamaru 1981). More recent analyses indicate the existence of 12 or more cryptic lineages within A. vulpes with sister lineages separated by at least 3.5 Ma (Bowen et al. 2007; Hidaka et al. 2008).

Colour is often used to delineate closely related taxa especially among the vibrantly coloured species that inhabit tropical coral reefs. However, discordance between genetic divergence and coloration is well documented in butterflyfishes (McMillan & Palumbi 1995) and hamlets (Ramon et al. 2003; Garcia-Machado et al. 2004). The brightly coloured pygmy angelfishes (Genus Centropyge) offer striking examples. The Flame Angelfish (Centropyge loricula) maintains distinct geographic colour morphs across its Central Pacific range without corresponding genetic partitioning at neutral loci (Schultz et al. 2007). Centropyge in the Atlantic have largely allopatric distributions and heritable colour patterns, yet they are phylogenetically indistinguishable (Bowen et al. 2006). Three species have been described in the Centropyge flavissima complex based on colour; however, distinct mtDNA lineages cluster by geographic location (Pacific Ocean, Indian Ocean and Moorea) and not by species (DiBattista et al. 2012). In the Blackstreak Cleaner Wrasse Labroides dimidiatus, mtDNA lineages group by Pacific archipelago and not colour morph (Sims et al. 2013).

Acanthurus reversus was described based on diagnostic colour differences: the orange band in A. reversus is much smaller than in A. olivaceus, and the tail has a ‘reverse’ coloration, with a white margin and a black submarginal band, whereas in A. olivaceus, the margin is black and the submarginal band is white (hence the name, Acanthurus reversus; Fig. 1a). Analyses of mtDNA and nuclear introns indicate that this colour change is not accompanied by reciprocal monophyly (Figs 1c and 2). Only with the analysis of the much larger RADSeq data set is A. reversus distinguished from its sister taxa with a strong signal of selection (Fig. 4). Coalescent analysis indicates that evolution of colour in A. reversus was rapid (TMRCA = 0.83 Ma) relative to the species complexes of Centropyge described above (TMRCA = 2.7–12 Ma, Gaither et al. 2014). Taken together, these studies confirm that colour evolves at different rates across groups, does not necessarily correspond with the evolutionary partitions detected in phylogenetic studies and should be used as a taxonomic trait with caution.

Selection on rhodopsin genes

Our RADSeq analysis revealed 59 loci (of 3737) under divergent selection, with one of them (locus 65 948; Table S4, Supporting information) mapping to the opsin Rh2 gene. Opsins are a group of light-sensitive pigments involved in mediating the conversion of light into an electrochemical signal, the first step in the visual transduction cascade. The Rh2 gene is an essential com-
ponent of the cone receptor system that is sensitive to middle wavelengths of the visible spectrum (Owens et al. 2009).

Coral reefs are generally characterized by clear oligotrophic waters (Birkeland 1988); however, coral reefs in the Marquesas are atypical. Episodic increases in nutrients around the islands, attributed to a combination of ocean currents, El Niño events, and terrestrial run-off, yield algal blooms and lower visibility (Randall & Earle 2000; Martinez & Maamaatuaiahutapu 2004). Selection in opsin genes has been detected in a number of species, including African cichlids, where fixation of different alleles has been shown to correspond to adaptation to different light regimes and coincides with divergence in male breeding coloration (Spady et al. 2005; Sugawara et al. 2005; Terai et al. 2006). Further, positive selection on Rh1 and Rh2 opsin genes may have been important in the early evolution of the group (Dann et al. 2004). In marine gobies (family Gobiidae), amino acid substitutions at spectral turning sites of the Rh1 gene correspond with divergent light environments related to depth, turbidity and light spectrum (Larmuseau et al. 2010). In the Cormoran coelacanths (Latimeria chalumnae), amino acid replacements at opsin genes have been tied to adaptation to the low light conditions at depth (Yokoyama et al. 1999). Taking these examples into account, the low light conditions that characterize the Marquesas could be driving the selection in visual pigments detected here in A. reversus. Future studies targeting opsin genes across taxonomic groups will help inform the role of these pigments in adaptation of fishes at the Marquesas.

The significant shifts in haplotype frequency detected in A. reversus, combined with a dramatic divergence in colour and signals of selection that correspond with known biogeographic partitions, indicate A. reversus is on an independent evolutionary trajectory. While the lack of monophyly may be unsatisfactory to some, we prefer to avoid unnecessary nomenclature juggling and therefore suggest no change in taxonomy.

The utility of RADSeq in phylogeographic studies

Over the past few decades, the number of mtDNA studies has grown rapidly and the inferences drawn from these data sets have greatly improved our understanding of phylogeography, evolution and ecology. From these data, we can calibrate molecular clocks across a diversity of marine taxa (Bowen et al. 2001; Lessios 2008; Reece et al. 2010). However, gene trees generated from the maternally inherited mtDNA genome, while biologically informative, may not represent the species evolutionary history. The alternative of producing nuclear intron sequence data can be time-consuming and expensive, which limits the number of loci (and the proportion of the genome) analysed. Moreover, sequencing approaches that target neutral loci are unlikely to reveal signals of selection.

With new and user-friendly laboratory protocols (Peterson et al. 2012; Toonen et al. 2013) and bioinformatics pipelines (Catchen et al. 2011; Puritz et al. 2014), as well as the falling price of sequencing, we can now resolve thousands of loci across the genomes of nonmodel organisms quickly and cost effectively. The utility of this approach, as well as others that target specific genomic regions (i.e. target enrichment or sequence capture), is changing the way we conduct phylogeographic studies (McCormack et al. 2013; Rocha et al. 2013; Andrews & Luikart 2014; Bowen et al. 2014). While the nuclear introns assayed here showed a near complete lack of genetic structure, the RADSeq data revealed low levels of population structure at Bali (Tables 3 and 4). The latter may be due to the position of Bali at the western end of the species range. Perhaps the most interesting aspect is, however, the ability to detect divergent selection at the Marquesas, a finding only made possible with the amplification of 1000s of loci randomly distributed across the genome.

Conclusion

Phylogenetic and phylogeographic analyses of neutral loci may lack the resolution to detect recent isolation or rapid divergence. The distinct ecological conditions at the Marquesas, and the finding of genetic divergence across numerous taxa, indicate that both selection and genetic drift are driving divergence at these islands. The detection of 59 outlier loci, combined with divergence in coloration and a signature of selection in visual pigment genes, indicates that A. reversus is on an independent evolutionary trajectory, driven by divergent selection in the unique environment at the Marquesas. The synchronous isolation in Hawai’i and the Marquesas demonstrates how adaptation can accelerate the speciation process. In the half million years since these populations were founded, the geographic isolation of Hawai’i has produced a divergent population, whereas the unique ecology of the Marquesas has produced what taxonomists recognize as a distinct species. This is the first case study of its kind for coral reef fishes, and we believe that further studies will confirm that speciation driven by natural selection is an important source of biodiversity in the sea.

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M.G., B.W.B. and L.A.R. designed the project and supervised field operations. M.G., M.A.B., B.W.B. and L.A.R wrote the initial manuscript. M.G., M.A.B., R.R.C. and S.A.J. produced Sanger data and prepared RADSeq libraries. W.B.S. performed quality control of RADSeq data and provided bioinformatics support. M.G., M.A.B. and R.R.C. analysed data. All authors contributed to editing and revising the manuscript.

Data accessibility
Sanger sequence alignments and RADSeq raw data were deposited in Dryad, entry doi: 10.5061/dryad.581f3.

Supporting information
Additional supporting information may be found in the online version of this article.

Fig. S1 Phylogenetic tree for cytochrome b sequences of Acanthurid species for which sequences (> 600bp) were publicly available (Table S2, Supporting information).

Table S1 Morphometrics for species of Acanthurus from Randall (2005).

Table S2 List of species and GenBank accession numbers for cytochrome b sequences used in phylogenetic analyses (Fig. S1, Supporting information).

Table S3 Pairwise FST values for eleven populations of Acanthurus olivaceus and A. reversus from the Marquesas.

Table S4 Blastn top hit for each of 59 outlier loci with a sequence match of higher than ≥ 50% and an E-value of ≤ e-10 or lower.

Appendix S1 Materials and methods.