1 INTRODUCTION

Long QTc on electrocardiogram (ECG) can lead to torsade de pointes (TdP). Although this is usually self-limiting, it may degenerate into ventricular fibrillation. Prevention or treatment of the factors underlying long QT would be expected to have significant clinical benefit. Hospitalised patients may be at greater risk of long QT than the general population. Acute medical admissions tend to be elderly...
and so may be more vulnerable for drug-induced QTc prolongation and thence arrhythmia, as the QTc interval is known to increase with age.\textsuperscript{2,3} In addition, elderly patients may have other risk factors for QTc interval prolongation, such as the presence of diabetes mellitus, polypharmacy or electrolyte disturbances.\textsuperscript{4}

QT prolongation has been shown to be present in up to 45\% of patients after cerebrovascular accident (CVA)\textsuperscript{5} and in 18\% of patients on a coronary care unit (CCU)\textsuperscript{6} but a much lower prevalence (0.9\%) was recorded in a single centre study across all inpatients of a general hospital.\textsuperscript{7} The incidence and prevalence of prolonged QTc in unselected medical admissions to an acute medical unit (AMU) is unknown and the frequency with which QT interval-prolonging drugs are prescribed to patients at admission and following admission, in this environment, are also unknown. We hypothesised that QT prolongation would be more frequently encountered in acute medical admissions than in a hospital-wide environment,\textsuperscript{7} thereby posing a considerable clinical risk.

The objectives of this study were to (a) determine the incidence of long QT upon admission to a large urban teaching hospital with an acute medical illness; and (b) determine the proportion of patients receiving QT-prolonging drugs at the time of admission and the change in prescribing of QT-prolonging drugs after admission.

2 | METHODS

A retrospective study of 1000 consecutive patients admitted to an AMU at a 950-bed, urban teaching hospital. At our institution, ST-elevation myocardial infarction and acute CVA are triaged directly to the cardiac and stroke services, respectively, and so are not referred to the AMU. Patients routinely have a 12-lead ECG within 12 hours of admission (Welch-Allyn\textsuperscript{9}). The ECG trace was accessed using EDMonline\textsuperscript{8} software. The first satisfactory ECG obtained for each ED visit was included. If the ECG was noisy, obscuring the relevant onset or offsets, incomplete with some leads missing, or rapid with overlap of the terminal T wave and the next cycle P or QRS waves, then the next ECG obtained during the same ED visit was assessed. If there was no satisfactory ECG, then the case was excluded. Other exclusions were patients aged less than 18 years of age, paced ECG, missing patient ECG or missing patients’ identifiers.

The QT interval represents the time from the beginning of ventricular depolarisation to completion of repolarisation. However, because the QT interval encompasses ventricular depolarisation adjustment of QTc in those with left bundle branch block was made using the formula: measured QT interval minus 50\% of LBBB duration.\textsuperscript{8,9}

The presence and severity of candidate risk factors (determined from the published literature) were recorded (obtained from the EPR system). These fell under three domains: basic demographic (age, gender\textsuperscript{10,11}), biochemistry (capillary glucose at the time of the ECG,\textsuperscript{12} potassium,\textsuperscript{13} calcium and magnesium concentrations\textsuperscript{10,14} obtained closest temporally to ECG acquisition within 12 hours) and medications prescribed within 7-days of the index ECG. Community-issued medications are obtained by pharmacist reconciliation with primary care. All data were anonymised. A medication was classified as QTc-prolonging based upon listing at a pharmacy-reference website (www.crediblemeds.org).\textsuperscript{15} In this listing, medications are sub-classified as ‘known’, probable’ or ‘conditional’ for QTc prolongation, or whether they are at risk of inducing QTc prolongation in those with congenital long-QT syndrome.

The principal outcome measure was the incidence of long QTc using Bazett formula in acute medical admissions. Bazett formula is

\[
\text{QTc} = \frac{\text{QT}}{(HR/60)^{1/2}}
\]

Secondary outcome measures were: proportion of patients with prolonged QTc who were subsequently administered drugs known to prolong QTc during their inpatient stay; length of admission; in-hospital mortality; rate of readmission at 1, 3 and 12 months; inter-observer agreement (kappa statistic) of QTc length. We also compared the incidence of QT based upon the automatic ECG output, which uses the Bazett formula by default and with the Fridericia formula (QTcF), for which the formula is

\[
\text{QTc} = \frac{\text{QT}}{(HR/60)^{1/3}}
\]

QTcF is thought to reflect a more accurate correction factor than Bazett’s formula in subjects with faster heart rates.\textsuperscript{17}

QT intervals were determined manually by two investigations. QT intervals were determined manually from lead II of 12-lead ECGs. QT intervals were measured from the earliest QRS deflection to the end of the T wave. The intersection of the down slope of the T wave with the baseline was taken as the end of the QT interval. With a sinus rhythm, QT and RR intervals were averaged over three consecutive complexes. During other rhythms, QT and RR intervals were averaged over all complexes on lead II. QTc of ≥500 milliseconds was considered significantly prolonged in both genders; >470 milliseconds in men and >480 milliseconds in women was considered moderately prolonged\textsuperscript{6,18} and represents the 99th percentile for QTc.\textsuperscript{19} Cohen’s kappa statistic,
for inter-observer agreement, was made from a subgroup of 200 randomly chosen ECGs evaluated by a third observer. Ethical permission was sought but was considered not necessary by the hospital’s Research and Innovation (R&I) department.

2.1 Statistic analysis

Normality of distribution was assessed using the Kolmogorov–Smirnov test. Comparison of normally distributed data was made by student’s unpaired t test assuming equal or unequal variance as necessary. Correlation of heart rate with Bazett formula and to Fridericia formula was with Pearson’s correlation. Comparison of length of stay was made with Wilcoxon rank sum test. The proportion of patients surviving their hospital admission, and proportions being readmitted within 1, 3 or 12 months after discharge was made by a 2 × 2 Fisher exact test, with Sidak correction for multiple testing.

A multivariate logistic regression model was used to determine variables relating to the presence/absence of long QTc. The variables decided upon, a priori, to be included were: age, gender, potassium, calcium, magnesium concentrations, medication use and presence of diabetes. To examine the influence of long QTc on length of stay, a multiple linear regression model was constructed that included any variable associated with long QTc with \( P < 0.10 \). SPSS 24.0 (SPSS Inc, Chicago, IL) was used for statistical analyses. Data are mean ± SD unless stated otherwise. Statistical significance was considered if \( P < 0.05 \).

3 RESULTS

Of 1000 consecutive admissions 285 cases were excluded, leaving 712 patients (361 males and 351 females; Figure 1). The reason for admission is provided in Table S1. There were 50 cases of QTc prolongation (31 males and 19 females) giving an incidence of 7%. Of the 50 cases, \( n = 12 \) were severe QTc prolongation (≥500 milliseconds), an overall incidence of 1.7%. Of the 50 cases, \( 18 (36\%) \) had bundle branch block. Of the 662 without long QT, 75 (11.3%) had bundle branch block (\( P < 0.001 \)).

The automatic QTc measurement recorded a significantly greater proportion of patients with long QTc (\( n = 111 \) [15.6%], \( P < 0.001 \)), and severe QTc prolongation (\( n = 42 \) [5.9%], \( P < 0.001 \)) than with manual Bazett. A sensitivity analysis was performed for the comparison of manual vs automatic classification of long QTc by excluding all patients with bundle branch block: 71 (11.5%) patients had long QTc by automatic reader (\( n = 548 \) with normal QTc), whereas 32 (5.2%) had long QTc by manual Bazett (\( n = 587 \) with normal QTc; \( P < 0.001 \)).

The heart rate was 87.0 ± 22.2 per minute. QTcF identified only nine patients (1.3%) with prolonged QT, of which three (0.4%) were ≥500 milliseconds. The mean (95% confidence interval) for the difference in QTc between Bazett and Fridericia formulae was 23.1 milliseconds (21.8–24.3). The Pearson’s correlation between heart rate and QTc with Bazett was \( r = 0.157 \) (\( P < 0.001 \)) and heart rate to QTcF was \( r = -0.270 \) (\( P < 0.001 \)).

Admission potassium was available in \( n = 581 \) (82%), calcium in \( n = 703 \) (99%), magnesium in \( n = 603 \) (85%). The admission
potassium and the proportion of patients using QT-prolonging drugs on admission were significantly different in those patients with prolonged QTc (Table 1).

Kappa inter-observer agreement for long QTc with manual Bazett was 0.65 indicating substantial agreement.

3.1 | In-hospital prescription of QTc-prolonging drugs

At admission, 328 patients (46%) were receiving ≥1 QTc-prolonging drug of any kind. Of this group, 96 patients were using medication known to cause TdP, 70 were using drugs with a possible risk of TdP, 231 using medications with a conditional risk of TdP and 13 were using medications with a risk in congenital long-QT syndrome.

Following admission, 512 patients had no change in the number of QTc-prolonging drugs, 72 had a reduction in number and 105 patients had an increase (Figure 2). Of those with prolonged QTc at admission, seven subsequently had a reduction in number of QT-prolonging medications, 37 were unchanged and six had an increase (Figure 3).

The QTc-prolonging drugs observed in this study are shown in Table 2. The mostly frequently prescribed drug (with a known risk of TdP) was Citalopram, although clarithromycin was the most commonly instituted drug in this risk group. Furosemide was the most commonly prescribed drug with any risk of QTc prolongation.

3.2 | Patient outcomes

Length of stay was greater in those with prolonged QTc (median 7.2, IQR 2.4-13.2 days) than in those without (median 3.3, IQR 1.3-10.0 days; \( P = 0.004 \)). However, in a linear regression model for length of stay, using predictor variables with \( P < 0.1 \) (from Table 1), only patient age was significant (\( \beta \) coefficient 0.215; \( P < 0.001 \)). None of the following affected the length of stay: presence of long QTc (\( \beta = 0.031; P = 0.457 \)), presence of diabetes (\( \beta = 0.044; P = 0.293 \)), admission potassium (\( \beta = -0.08; P = 0.058 \)). Twenty-eight patients died during the index admission (age 84 [IQR 77.5-88.5] years vs 65 [IQR 49-79]; \( P < 0.001 \)). Three individuals who died had long QTc, of which one (solid-organ cancer) was severely prolonged. There were no differences in the proportions of those who survived admission or in the rate of readmission (Table 3).

4 | DISCUSSION

This study has shown that approximately 1 in 14 patients admitted to an acute unit had prolonged QT interval as judged by a manual Bazett method. Ventricular arrhythmias are most often associated with QTc values of 500 milliseconds or more\(^{17,20}\) and in this group (severe long QTc) the incidence was 1.7%. This is the first major study reporting incidence in acute medical admissions.

| Patient characteristic       | Normal QTc \( n = 662 \) | Long QTc \( n = 50 \) | \( P \) value |
|------------------------------|--------------------------|----------------------|--------------|
| Age                          | 62.8 ± 19.3              | 68.1 ± 19.7          | 0.060        |
| Gender                       | Male = 330; female = 332 | Male = 31; female = 19 | 0.108       |
| Presence of diabetes mellitus| Yes = 176; no = 486      | Yes = 19; no = 31    | 0.099        |
| Serum potassium at admission (mmol/L) | 4.3 ± 0.7               | 3.9 ± 0.7            | <0.001       |
| Serum calcium at admission (mmol/L) | 2.25 ± 0.12             | 2.24 ± 0.13          | 0.564        |
| Serum magnesium at admission (mmol/L) | 0.78 ± 0.11             | 0.76 ± 0.18          | 0.554        |
| Presence of QTc-prolonging drugs at admission | Yes = 294; no = 368 | Yes = 35; no = 15 | 0.009 |

TABLE 1 Comparison of candidate factors reported to affect QTc

Long QTc defined as >470 ms males; >480 ms females.
By contrast, prevalence data have shown that prolonged QTc is frequently encountered in an acute stroke unit or CCU. These data likely reflect the case mix of the various cohorts, as transmural myocardial ischaemia and intraventricular haemorrhage are factors associated with long QT, but are usually seen within a high-dependency environment. However, long QT (particularly drug-induced long QT) has also been associated with increasing age, congestive cardiac failure, hypokalaemia and anti-psychotic medication. Individuals with major psychiatric disorders, with cardiovascular diseases and elderly people comprise a sizeable proportion of the acute medical population. Acutely ill patients are likely to have multiple acquired risk factors for QT prolongation. Our data suggest that low-normal plasma potassium has a role in prolonging the QT interval and is a concern given the large number of patients prescribed potassium wasting diuretics. Hypokalaemia (or low-normal plasma potassium) should be sought and corrected following the initial clinical encounter, especially in those with (or at risk) of QTc prolongation.

**FIGURE 3** Change in number of medications with potential to prolong QT interval, following admission, in patients with pre-existing long QTc

**TABLE 2** Prescription of QT-prolonging drugs following admission

| Known risk of torsades de pointes | Possible risk of torsades de pointes | Conditional risk of torsades de pointes | Risk in congenital LQTS |
|----------------------------------|----------------------------------|---------------------------------|----------------------|
| Citalopram n = 34 (+1)           | Mirtazapine n = 27 (+1)          | Frusemide n = 86 (−8)           | Budesonide/Formoterol n = 7 (+2) |
| Quinine n = 22 (+2)              | Risperidone n = 15 (+2)          | Amitriptyline n = 28 (−1)      | Co-trimoxazole n = 2 (+2) |
| Clarithromycin n = 20 (+15)      | Ciprofloxacin n = 14 (+8)       | Sertraline n = 25 (−3)         | Salmeterol n = 1 (−1) |
| Domperidone n = 11 (+2)          | Promethazine n = 20 (0)         | Metoclopramide n = 24 (+6)     | Terbutaline n = 1 (−1) |
| Ondansetron n = 11 (+6)          | Venlafaxine n = 5 (−1)          | Omeprazole n = 15 (+3)         | Formoterol n = 0 (−1) |
| Haloperidol n = 10 (+6)          | Buprenorphine n = 4 (+1)        | Metronidazole n = 11 (+10)     |                      |
| Amiodarone n = 9 (+4)            | Aripiprazole n = 3 (0)          | Fluoxetine n = 9 (−4)          |                      |
| Donepezil n = 7 (0)              | Nortriptyline n = 2 (0)         | Indapamide n = 9 (−3)          |                      |
| Methadone n = 6 (0)              | Alfuzosin n = 1 (0)              | Olanzapine n = 9 (+4)          |                      |
| Fluconazole n = 3 (−1)           | Atomoxetine n = 1 (+1)          | Solifenacin n = 7 (−1)         |                      |
| Azithromycin n = 2 (0)           | Lithium n = 1 (0)               | Quetiapine n = 6 (−1)          |                      |
| Erythromycin n = 2 (−1)          | Morphine n = 1 (+1)             | Ritonavir n = 5 (0)            |                      |
| Flecainide n = 1 (0)             | Tolterodine n = 1 (0)           | Amisulpride n = 4 (0)          |                      |
| Levomepromazine n = 1 (0)        | Vardenafil n = 1 (0)            | Bendroflumethiazide n = 4 (−2) |                      |
| Sotalol n = 1 (0)                | Clozapine n = 0 (−2)            | Ivabradine n = 3 (0)           |                      |
| Dasatinib n = 0 (−1)             | Hydrochlorothiazide n = 2 (−2)  |                  |                      |
| Dextroamphetamine n = 0 (−1)    | Loperamide n = 2 (−3)           |                  |                      |
| Efavirenz n = 0 (−1)             | Paroxetine n = 2 (0)            |                  |                      |
| Chloroquine n = 1 (+1)           |                               |                  |                      |
| Posaconazole n = 1 (0)           |                               |                  |                      |
| Ranolazine n = 1 (−1)            |                               |                  |                      |
| Trazodone n = 1 (−2)             |                               |                  |                      |
| Voriconazole n = 1 (−1)          |                               |                  |                      |
| Hydroxyzine n = 0 (−2)           |                               |                  |                      |

Changes in prescribing are in parentheses.
TABLE 3  Clinical outcomes of long QT

| Outcome                              | Normal QTc N = 662 | Long QTc N = 50 | P value  |
|--------------------------------------|-------------------|-----------------|----------|
| Length of stay (days); median (IQR)  | 3.3 (1.3-10.0)    | 7.2 (2.4-13.2)  | 0.004    |
| In-hospital mortality                |                   |                 |          |
| Yes = 25 (3.8%); no = 637            |                   |                 | 0.438    |
| Yes = 3 (6.0%); no = 47              |                   |                 |          |
| 1-month readmission (excluding in-hospital deaths) | Yes = 91; no = 544 | Yes = 10; no = 37 | 0.202 |
| 3-month readmission                  |                   |                 |          |
| Yes = 179; no = 456                  |                   |                 | 0.405    |
| 12-month readmission                 |                   |                 |          |
| Yes = 304; no = 331                  |                   |                 | 0.366    |

Automatic QT interval measurement, integrated into the electrocardiographic device, is convenient for routine use but can be imprecise. In our series, the automatic measure overestimated the QT interval compared to manual evaluation—even in patients without bundle branch block. Why this occurred is uncertain although discordance has been suggested to relate to difficulties in delineating the end of T wave either due to ‘noise’ or in cases where the T wave is flat, bifid, biphasic, or overlapping on a U wave. Conversely, the Fridericia formula gave a far lower estimation of incidence of prolonged QT, with a mean difference of 23 milliseconds between Bazett and Fridericia formulae. This difference was larger than previously reported. Fridericia formula may be more accurate in the context of tachycardia, although Bazett's formula is thought to adequately correct for heart rates varying between 50 and 90 beats per minute, of note the mean heart rate in this study was 87/min and 40% of patients had a pulse over 90/min. The correlation data suggested that QTc tended to shorten with tachycardia using Fridericia, whereas Bazett formula had a positive correlation with heart rate.

The prescribing patterns reflect the case mix seen on an AMU. NICE guideline recommends consideration of dual antibiotic therapy with amoxicillin and a macrolide (such as clarithromycin) for patients with moderate- or high-severity community-acquired pneumonia. Use of fluoroquinolones, which have been a cause for concern to their risk for TdP, are not promoted as first line. Safety issues have also been raised for use of QTc-prolonging anti-psychotic medications. Such drugs are frequently used in hospital inpatients as delirium can affect a fifth of acute medical admissions. NICE guidelines allow for the antipsychotics haloperidol and olanzapine to be used for up to 1-week as a therapy for delirium. Whether QTc contributes to increased hospital mortality in those with delirium is unknown. Prolonged QT interval is known to predict mortality in a variety of other conditions that are frequently encountered in the AMU such as: coronary artery disease, heart failure and diabetes mellitus. Survival curves of those with/without prolonged QTc separated well within 50 days of admission, in a hospital-wide study. Our data have shown in-hospital mortality of 3.8% in those with normal QTc and 6.0% of those with long QTc. The study was not powered for this outcome but such a difference would be clinically relevant; based upon these data, approximately 1500 patients would be needed to be studied to detect a difference with 80% power and alpha level of 5%. The more prolonged length of stay (in patients with long QTc) in the unadjusted analysis was unexpected but was no longer significant when correcting for patient age and presence of diabetes or electrolyte disturbance.

The strengths of our study include the robust inter-observer reproducibility of long QTc diagnosis, the high-degree of capture of biochemical data, and complete access to electronic drug prescriptions and pharmacy reconciliation with primary care. There are some limitations to our study. First, retrospective cohort studies are susceptible to selection bias although this was minimised using consecutive cases. Second, as discussed, the adequacy of Bazett’s formula has been questioned as it may overcorrect the QT interval at fast heart rates and under-corrects it at low heart rates. Third, we do not have data of the prevalence of coronary insufficiency or left ventricular dysfunction; these are risk factors for ventricular arrhythmia in the setting of prolonged QTc. Finally, this study was designed to determine the incidence of long QTc on presentation to an AMU; therefore the first ECG was evaluated but not any follow-up ECGs. It may be the case that a patient's long QTc resolved prior to use of QTc-prolonging drugs. However, prolonged QTc may also develop after admission. In a study of 178 patients prescribed QTc-prolonging drugs, 26 patients (15%) went on to develop a prolonged QTc within 11-days of starting a QTc-prolonging drug. Finally, we did not find that prolonged QTc independently affected length of stay or in-hospital mortality. This may represent a type II error as the study was not powered for these outcomes and the number of deaths was relatively low. Data from a HDU/ICU environment have suggested that, in an adjusted model, QT prolongation (captured on continuous bedside monitoring) nearly triples the odds for in-hospital mortality and prolonged length of stay.

In conclusion, we have shown that QT prolongation, judged by the Bazett formula, is relatively common in acute medical admissions and that low plasma potassium (albeit within the normal range) and QT-prolonging drugs contribute to this. The incidence of prolonged QTc varied considerably by the formula used for heart rate correction. Whether prolonged QTc interval affects clinical outcomes, in this group of patients, requires further study.

DISCLOSURE
R.M, A.G., A.N. and M.B.W. no conflict of interest.

AUTHOR CONTRIBUTIONS
R.M. and A.G. collected the clinical data and ECG analysis. A.N. independently reviewed the ECG data. R.M, A.G and M.B.W. analysed all the data and wrote the manuscript. All authors reviewed the
manuscript. M.B.W is the guarantor of this work and, as such, had full access to all the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis.

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**SUPPORTING INFORMATION**
Additional supporting information may be found online in the Supporting Information section at the end of the article.

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