Predictive and spatial analysis for estimating the impact of sociodemographic factors on contraceptive use among women living with HIV/AIDS (WLWHA) in Kenya: Implications for policies and practice

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ABSTRACT

Background Despite the availability and knowledge of various contraceptive methods, consistent utilisation in women living with HIV/AIDS (WLWHA) within the reproductive age group remains below the Sustainable Development Goals (SDGs) and Family Planning 2020 goals. This study examines the association between sociodemographic factors and contraceptive use including the effect of clustering tendencies of these factors on contraceptive usage among WLWHA in Kenya.

Methods Weighted multivariate logistic regression models were conducted to determine the association of sociodemographic factors on contraception use among WLWHA using the 2008–2009 Kenya Demographic Health Survey. Spatial autocorrelation techniques were used to explore clustering tendencies of these factors on contraception utilisation. Our study population included 304 HIV positive women, aged 15–49 years.

Results Among 304 HIV-positive women in our study population, 92 (30.3%) reported using one method of contraception. Contraceptive use was significantly associated with wealth and education after adjustment for other sociodemographic variables. Women classified as having low and middle wealth index were less likely to use contraceptives compared with women classified as having high wealth index. Similarly, women with primary education only were less likely to use contraceptives compared with women with secondary or higher education. Spatial autocorrelation revealed significant positive clusters with weak clustering tendencies of non-contraceptive use among different levels of wealth index and education within different regions of Kenya.

Conclusion These findings underscores the need for intervention programmes to further target socially disadvantaged WLWHA, which is necessary for achieving the SDGs.

INTRODUCTION

HIV/AIDS is a pandemic disease with devastating socioeconomic effects on the world’s population. Globally, about 35 million people were reported to be living with HIV/AIDS (PLWHA) in 2013. Sub-Saharan Africa accounts for 70% of the HIV burden worldwide.1,2 In East Africa, Kenya has the highest HIV burden with a prevalence rate of 6%, and majority of these cases have been attributed to unsafe sexual practices.1,3 Kenya has a record of 1.5 million PLWHA with 36 000 AIDS-related deaths in 2015; in addition, women within the reproductive age group of 15–49 years account for more than 50% of cases.3–5 The number of AIDS-related deaths in Kenya have progressively reduced by 32% from 2005 to 2013 due to improved access to antiretroviral therapy.1 Despite the advances in the
management of HIV/AIDS, substantial gaps exist in relation to the reproductive needs of women living with HIV/AIDS (WLWHA), particularly in accessibility and use of contraceptives.4

From a public health perspective, contraceptive use in WLWHA can directly impact a range of health conditions including prevention of unplanned pregnancies and abortions, prevention of mother-to-child transmission (PMTCT), reduction in sexually transmitted infections, decline in maternal deaths during pregnancies, reduction of HIV superinfection (infection with a different strain of HIV) and reduction in chances of HIV transmission to their uninfected partner.6 Therefore, addressing the reproductive health needs of WLWHA is critical to prevent unplanned pregnancy and to promote maternal and child health, a component of PMTCT action point as stated by the WHO, which has been shown to be more cost-effective than treatment of HIV positive mothers and babies with antiretroviral therapy.2

The Spectrum and Estimation and Projection Package (EPP) has been used in Kenya to model adult HIV infection rate, maternal-to-child transmission, child survival and maternal AIDS deaths in determining HIV prevalence. Despite this important contribution, there are gaps in the outputs generated from EPP models regarding plausible global incidence data and uncertainty in key assumptions employed in the model.10 Whereas the EPP model can project trends of HIV indicators and serve as a potent biosurveillance tool to monitor epidemic, it does not possess geographic information systems (GIS) mapping capabilities for determining HIV indices-related clusters.

Given the limited resources in Kenya, GIS is a valuable tool in providing insights into proximity analysis, disease clustering tendencies, identification of hotspots and performing spatiotemporal regression analysis that can improve the efficiency of distributing scarce resources to problematic hotspots. In a similar realm, some studies have employed the application of GIS and remote sensing techniques on georeferenced, normalised and spatiotemporal HIV-related datasets in analysing proximity geographic factors and contraceptive use among WLWHA, in particular, WLWHA in Kenya.

Therefore, the aims of this study were to: (1) investigate and quantify the relationship between sociodemographic factors and contraceptive use among WLWHA in Kenya and (2) create a robust model to identify clustering tendencies of sociodemographic factors on contraceptive use among WLWHA using spatial autocorrelation techniques, hotspot and cluster analysis.

### METHODS

#### Study design and data source

This cross-sectional study used data from the 2008–2009 Kenya Demographic Health Survey (KDHS).17

| Variables | Total (n=304) | Using (n=92) | Not using (n=212) | P value |
|-----------|--------------|--------------|------------------|---------|
| Region    |              |              |                  |         |
| Nairobi   | 36 (10.8)    | 14 (17.9)    | 22 (7.9)         | 0.2088  |
| Coast     | 29 (5.4)     | 9 (3.3)      | 20 (6.2)         |         |
| Nyanza    | 118 (35.3)   | 34 (35.7)    | 84 (35.1)        |         |
| Rift Valley | 29 (20.0)   | 9 (15.8)     | 20 (21.7)        |         |
| Western   | 46 (12.5)    | 9 (7.8)      | 37 (14.4)        |         |
| Others    | 46 (15.9)    | 17 (19.5)    | 29 (14.5)        |         |
| Residence |              |              |                  |         |
| Urban     | 112 (36.9)   | 42 (39.8)    | 70 (25.9)        | 0.132   |
| Rural     | 192 (70.4)   | 50 (61.1)    | 142 (74.1)       |         |
| Ethnicity |              |              |                  |         |
| Kikuyu    | 38 (13.9)    | 17 (18.4)    | 21 (12.0)        | 0.2536  |
| Luhya     | 57 (23.3)    | 12 (14.5)    | 45 (26.8)        |         |
| Luo       | 131 (43.8)   | 40 (45.3)    | 91 (36.3)        |         |
| Others    | 78 (26.3)    | 23 (21.1)    | 55 (24.8)        |         |
| Highest educational level |              |              |                  | 0.0033  |
| No education | 22 (5.9)   | 4 (2.8)      | 18 (7.1)         |         |
| Primary   | 191 (64.1)   | 50 (50.6)    | 141 (69.4)       |         |
| Secondary or higher | 91 (30.0) | 38 (46.6) | 53 (25.3) |         |
| Wealth index |              |              |                  | 0.0003  |
| Low       | 97 (32.7)    | 14 (15.5)    | 83 (39.5)        |         |
| Middle    | 48 (16.8)    | 11 (11.1)    | 37 (16.2)        |         |
| High      | 159 (52.6)   | 67 (73.4)    | 92 (44.3)        |         |
| Religion  |              |              |                  | 0.745   |
| Roman Catholic | 76 (25.2) | 23 (23.8) | 53 (22.2) |         |
| Protestant/other Christian | 210 (71.9) | 63 (72.5) | 147 (71.6) |         |
| Others    | 18 (5.5)     | 6 (3.7)      | 12 (6.2)         |         |
| Age (years) |              |              |                  |         |
| 15–19     | 24 (7.9)     | 3 (2.9)      | 21 (8.7)         | 0.3343  |
| 20–24     | 54 (16.4)    | 11 (14.7)    | 43 (17.1)        |         |
| 25–29     | 63 (22.9)    | 23 (21.5)    | 40 (23.4)        |         |
| 30–34     | 62 (19.3)    | 22 (26.9)    | 40 (16.3)        |         |
| 35–39     | 45 (15.3)    | 16 (13.8)    | 29 (10.3)        |         |
| 40–44     | 31 (10.2)    | 9 (11.2)     | 22 (18.1)        |         |
| 45–49     | 25 (6.9)     | 8 (9.0)      | 17 (6.1)         |         |
| Marital status |           |              |                  |         |
| Married   | 143 (47.6)   | 44 (51.8)    | 99 (43.2)        | 0.2967  |
| Not married | 161 (54.4) | 48 (48.2) | 113 (56.8) |         |

Counts and percentages represent unweighted frequencies and weighted percentages of study population. P value is derived from the Rao-Scott \( \chi^2 \) test. Percentages may not sum to 100 due to rounding.

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The Kenyan DHS is a nationally representative survey conducted every 5 years by Kenyan health agencies to assess the country’s health status. The KDHS employed a stratified two-stage sample design by identifying 400 sampling units or clusters as identified by the Kenyan Population census and sampled households from all eligible households within each cluster. Respondents to the survey included men and women aged 15–54 years and 15–49 years, respectively. Data collected include indicators of fertility, reproductive health, maternal and child health, mortality, nutrition and self-reported health behaviours among adults.

We obtained data on three domains: demographics and general health (DGH), Global Positioning System (GPS) and deidentified HIV data to protect the privacy of participants. Data on HIV status was obtained after a written request to the DHS division of the USAID was granted. Records from the HIV and DGH datasets were linked by cluster, house number and line number. These were subsequently linked with the GPS data by cluster and region to compile the final data for analysis.

**Study population**
The source population consisted of 8444 women. Our initial study population included 318 women aged

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**Figure 1** Prevalence map of WLWA aged 15–49 years per 100,000. WLWA, women living with HIV/AIDS.
Table 2  Regression analysis showing ORs and 95% CIs for the association of sociodemographic factors and contraceptive use among HIV-positive women in Kenya

| Variables                  | COR (95% CI)       | AOR (95% CI)       |
|----------------------------|--------------------|--------------------|
| Region                     |                    |                    |
| Nairobi                    | 2.22 (0.92 to 5.36) | 0.74 (0.17 to 3.13) |
| Coast                      | 0.52 (0.17 to 1.61) | 0.27 (0.05 to 1.55) |
| Rift Valley                | 0.71 (0.18 to 2.80) | 0.51 (0.08 to 3.47) |
| Others                     | 1.32 (0.58 to 3.03) | 1.16 (0.29 to 4.70) |
| Western                    | 0.53 (0.21 to 1.31) | 0.80 (0.20 to 3.20) |
| Nyanza                     | Reference (1.0)    | Reference (1.0)    |
| Residence                  |                    |                    |
| Rural                      | 0.55 (0.26 to 1.19) | 1.04 (0.36 to 3.01) |
| Urban                      | Reference (1.0)    | Reference (1.0)    |
| Ethnicity                  |                    |                    |
| Kikuyu                     | 1.22 (0.42 to 3.59) | 0.68 (0.14 to 3.27) |
| Luhya                      | 0.44 (0.15 to 1.26) | 0.33 (0.06 to 1.75) |
| Others                     | 0.71 (0.35 to 1.45) | 0.51 (0.14 to 1.87) |
| Luo                        | Reference (1.0)    | Reference (1.0)    |
| Highest educational level  |                    |                    |
| No education               | 0.20 (0.05 to 0.75) | 0.60 (0.12 to 2.98) |
| Primary                    | 0.37 (0.17 to 0.81) | 0.42 (0.18 to 0.98) |
| Secondary or higher        | Reference (1.0)    | Reference (1.0)    |
| Wealth index               |                    |                    |
| Low                        | 0.24 (0.11 to 0.52) | 0.17 (0.07 to 0.43) |
| Middle                     | 0.41 (0.17 to 1.01) | 0.33 (0.11 to 0.98) |
| High                       | Reference (1.0)    | Reference (1.0)    |
| Religion                   |                    |                    |
| Protestant/other Christian | 0.94 (0.47 to 1.91) | 1.06 (0.48 to 2.37) |
| Others                     | 0.55 (0.12 to 2.47) | 1.11 (0.16 to 7.75) |
| Roman Catholic             | Reference (1.0)    | Reference (1.0)    |
| Age*                       | 1.08 (0.9 to 1.3)  | 1.22 (0.99 to 1.51) |
| Marital status             |                    |                    |
| Not married                | 0.71 (0.37 to 1.4)  | 0.57 (0.29 to 1.15) |
| Married                    | Reference (1.0)    | Reference (1.0)    |

The AOR models include adjustment for age, marital status, region, residence, religion, ethnicity, education attainment and wealth index.

*Age was included as an ordinal variable.

AOR, adjusted OR; COR, crude OR.

15–49 years with a documented HIV positive result as determined by a laboratory test on dried blood spot samples conducted by the KDHS. We excluded women with incomplete data on study variables (n=14). The final study population consisted of 304 HIV-positive women.

Ethical approval
In addition, we obtained permission by written request and subsequent approval from the DHS division of the USAID.

Patient and public involvement (PPI)
For this analysis, there was no PPI involvement; however, for the original project from which data was obtained, PPI participation was essential.

Study variables
The outcome of interest was contraceptive use, and subjects were classified as using or not using contraceptives. The independent variables were sociodemographic characteristics. These sociodemographic characteristics included age in 5-year groups (15–19, 20–24, 25–29, 30–34, 35–39, 40–44 and 45–49 years), marital status (married or not married), region (Nairobi, Coast, Nyanza, Rift Valley, Western and others: North Eastern, Eastern and Central), residence (urban or rural), ethnicity (Kikuyu, Luhya, Luo or others: Embu, Kalenjin, Kamba, Kisii and Masai), highest educational level (no education, primary, secondary or higher), wealth index (low, middle or high) and religion (Roman Catholic, Protestant/other Christian denominations and other religious groups).

GPS coordinates for each respondent were included in our analysis to graphically represent the clustering of subjects. The DHS randomly displaced these coordinates by 2 km in the urban and 5 km in the rural zones, with 1% of the rural clusters further displaced by 10 km to protect the confidentiality of respondents.

Statistical analysis
Analysis was conducted using procedures for complex survey design, and sample weights were applied to all analysis. The characteristics of the study groups were described using unweighted frequencies and weighted percentages. Differences in the proportion of sociodemographic characteristics between contraceptive users and non-users were tested using Rao-Scott $\chi^2$ test. Crude ORs were computed by examining each sociodemographic variable with contraceptive use using weighted simple logistic regression models. All sociodemographic variables, age, marital status, region, residence, religion, ethnicity, education attainment and wealth index, were included in a weighted multiple logistic regression model to determine the sociodemographic factors independently associated with contraceptive use. Given our sample size and numerous categories within age, we included age as an ordinal variable in our multiple logistic regression model. ORs and their corresponding 95% CIs were reported.

Spatial analysis
A prevalence map was constructed on the DHS dataset to illustrate the distribution of WLWHA in Kenya in ArcMap. The prevalence was calculated by dividing the numerator (WLWHA) by denominator (total population per region). Hotspot Analysis (Getis-Ord Gi*) in
ArcGIS was employed to identify geographic landscapes that have significant high clusters and areas with significant low clusters based on three levels of CI. The z-score and p values parameters are indicators of statistical significance that enables the analyst to decide whether to accept or reject the null hypothesis. They also indicate whether the observed spatial clustering of high or low values is more pronounced than one would expect in a random distribution of those same values. A high z-score and small p value (p < 0.05) for a feature indicates a spatial clustering of high values. A low negative z-score and small p value indicates a spatial clustering of low values. The higher the z-score, the more intense the clustering.

A z-score near zero indicates no apparent spatial clustering. The input feature class for the hotspot analysis was the non-contraceptive layer, while the fixed distance band allowed for each centroid to have a neighbour. However, Euclidean intersite distance method was used to measure distance between two straight georeferenced clusters.

High and low clustering tool (Getis-Ord General G) is an inferential statistic that is interpreted in the context of the null hypothesis. The null hypothesis for the General G statistics (no spatial clustering) is rejected when the z-score is very high and the p value is low. Similarly, when the z-score is positive, it is interpreted as high values.

Figure 2  Hotspot analysis of contraceptive non-use among WLWHA in Kenya per region. WLWHA, women living with HIV/AIDS.
clustering within a georeferenced neighbourhood, and if the z-score is negative, it can be interpreted as clusters of low values within a study site.\textsuperscript{18–21} The methodology involves opening the Arc Toolbox, clicking on the spatial statistic tool and analysing pattern in ArcMap.

Also, the cluster and outlier tools were employed to identify cluster of features with similar values in magnitude and spatial outliers by creating a statistical algorithm for calculating a local Moran’s I value, a z-score, a p value and a code representing the cluster type for each statistically significant feature. Large z-scores signify an intense clustering of high values (HH) and a low z-score demonstrates more intense clustering of low values (LLs).\textsuperscript{18–20} The cluster/outlier type field in Anselin’s local Moran’s Index distinguishes between a statistically significant cluster of high values (HH), cluster of low values (LL), outlier in which a high value is surrounded primarily by low values (HL), and outlier in which a low value is surrounded primarily by high values (LH).\textsuperscript{22} Statistical significance was set at the 95\% CI, and no false discovery rate correction was applied to influence the critical p values.\textsuperscript{19} Variables with p values less than 0.05 were considered statistically significant.

**RESULTS**

The study population consisted of 304 HIV positive females, aged between 15 years and 49 years. Of these, 92
Figure 4  Cluster analysis of education on contraceptive non-use among WLWHA in Kenya. WLWHA, women living with HIV/AIDS.
Figure 5  Hotspot analysis of education on contraceptive non-use among WLWHA in Kenya. WLWHA, women living with HIV/AIDS.
(30.3%) reported using contraceptives (table 1). About 30% of the study population resided in urban areas, while 70% lived in the rural areas. Approximately 46% of women were married, and 46% of women were between 15 years and 29 years. Furthermore, 10.8% were from the Nairobi region, 5.4% from the Coast, 33.3% from Nyanza, 20.0% from Rift Valley, 12.5% from Western area and 15.9% from other regions. Most women were from the Luo ethnic group (38.9%), followed by Luhy (23.3%) and Kikuyu (13.9%); approximately 17% of study subjects were from other ethnic groups. A majority of women were classified as high-income earners (52.6%), while 14.8% and 32.7% were classified as middle-income and low-income earners, respectively. Approximately 5.9% of these women had no formal education, 64.1% attained primary education and 30% attained secondary or higher education. With respect to religion, 72% of women reported as Protestants or from other Christian denominations, 22.6% were Catholics and 5.5% were from other religions.

Table 1 shows the distribution of sociodemographic characteristics across study groups. There were no statistically significant differences in the proportion of these characteristics except for educational attainment and wealth index. Among women who reported not using contraceptives, 7.1% had no formal education compared with 2.8% of women who used contraceptives. Conversely, women who used contraceptives were more educated than women who did not use any form of contraception (46.6% vs 23.5%). These differences were statistically significant (p=0.0033). Furthermore, approximately 39% of women who do not use contraceptives have a low wealth index compared with 15% of women who use contraceptives (p=0.0033).

Figure 1 is a prevalence map illustrating the distribution of WLWHA in different regions of Kenya per 100,000 populations. Nyanza was observed to have the highest prevalence of women between the ages of 15–49 years living with HIV/AIDS (2.41–4.20/100 000 population) compared with Eastern, North Eastern and Rift Valley regions (0.20–0.60/100 000 population).

Table 2 reports the crude and adjusted ORs of the association between sociodemographic factors and contraceptive use among Kenyan WLWHA. Consistent with results from table 1, level of education and wealth index remained as independent predictors of contraceptive use in our adjusted models.

In univariate analyses, women without any primary education had 80% lower odds of using contraceptives compared with women with secondary or tertiary education (OR=0.2, 95% CI 0.05 to 0.75). However, this effect was substantially attenuated after controlling for other variables, including wealth index (OR=0.6, 95% CI 0.12 to 2.98). Women with primary education only were approximately 58% less likely to use contraceptives (OR=0.42, 95% CI 0.18 to 0.98) compared with women with secondary or higher education after adjusting for other sociodemographic characteristics.

Similarly, in univariate analyses, women with low wealth index were approximately 76% less likely to use contraception (OR=0.24, 95% CI 0.11 to 0.52) compared with women with high wealth index. This association persisted and increased slightly in magnitude after controlling for other variables (OR=0.17, 95% CI 0.06 to 0.43). In a dose–response fashion, women classified as having a middle wealth index were 59% less likely to use contraceptives compared with women classified as having a high wealth index (OR=0.41, 95% CI 0.17 to 1.02). After adjusting for other sociodemographic variables, the association increased slightly in magnitude and was marginally statistically significant (OR=0.33, 95% CI 0.11 to 0.98).

Getis and Ord Gi* analysis in figure 2 revealed hotspots for contraceptive non-use among WLWHA in Nyanza and Nairobi regions, while Rift Valley is a region with significant cold spots. Cluster analysis in figure 3 revealed high clusters of non-contraceptive use in WLWHA within localities of Nyanza and parts of Nairobi, whereas Rift Valley has low clusters of non-contraceptive use. Nevertheless, spatial outliers were important epidemiographic features seen for Nyanza and Nairobi. Furthermore, figure 4 revealed high spatial clustering of contraceptive non-use among study participants with primary education in Nyanza, and mainly cluster outliers for other levels of education in different regions of Kenya. Figure 5 further demonstrates that Nyanza and Nairobi regions show significant hotspot for education and study participants had at least primary form of education.

Table 3 revealed a moderately high z-score >2.56 and small p values <0.05 for wealth index (high, middle and low) and educational level (no education, primary and secondary or higher). Therefore, we rejected the null hypothesis that there was no spatial clustering for the predictors of interest. However, the clustering tendencies for these variables as indicated by the Moran’s Index is weak (<0.30) for contraceptive non-use among WLWHA in Kenya.
The objective of this study was to examine the relationship between contraceptive use and sociodemographic factors including the effect of clustering tendencies of these factors on contraceptive utilisation among WLWHA using spatial autocorrelation statistics. Findings from this study reveals wealth index and educational level to be significant sociodemographic predictors of contraceptive use. Women with low and middle wealth index were less likely to use contraceptives compared with women with high wealth index. This finding is consistent with previous studies in Mozambican and Malawian women where a lower use of modern contraceptives was found among women with low income. Decreased use of

**Figure 6** Hotspot analysis of Income on contraceptive non-use among WLWHA in Kenya. WLWHA, women living with HIV/AIDS.
Figure 7  Cluster analysis of income on contraceptive non-use among WLWA in Kenya. WLWA, women living with HIV/AIDS.
contraceptives observed for these women might suggest lack of accessibility and affordability of contraceptives in relation to those with high income.

From our results, 74% of non-contraceptive users live in rural areas compared with 61% of contraceptive users in our study. Given that people who reside in poorer communities generally have a slower uptake of good health practices such as contraceptive use, it is likely that the environmental influences on women with low wealth index might limit opportunities to access or afford contraceptives. This highlights a possible target for intervention by the Government of Kenya (GoK) through the provision of free or subsidised contraceptives for WLWHA, especially those living in rural areas with limited health facilities. Additionally, the Kenya Household Health Expenditure and Utilization Survey 2013 report stated that approximately 3% of Kenyans in the lowest wealth category have health insurance. Although the GoK waives the fee of services at government-owned hospitals, pharmacies and maternal care clinics, no programme explicitly targets WLWHA through the provision of low cost or free health insurance. Consequently, the financial responsibility of purchasing prescription drugs and modern contraceptives falls on the patient that could ultimately be a barrier to access. Therefore, efforts to improve insurance coverage for WLWHA could improve accessibility by reducing the financial burden associated with purchasing effective modern contraceptives.

Another important finding was that women with primary education only had a lower likelihood of contraceptive use than those with secondary and higher education. Similar findings were reported in previous studies highlighting a greater likelihood of using contraception in women with secondary education or higher. This might not necessarily translate to lack of contraception awareness as studies have shown knowledge of contraception among women across all educational levels to be nearly universal and yet does not imply usage. Thus, our findings may reflect the association of lower educational attainment with established social determinants of health such as poverty and unemployment that could impede access to primary health facilities and subsequent contraceptive use. Furthermore, the influence of male decision making and power dynamics in couple relationships could also explain the association between low educational level and contraceptive use among WLWHA. A recent qualitative study of married men in the Nyanza province of Kenya found that men may disapprove the use of contraceptives by their spouses due to fear of losing their position of authority in the relationship and promotion of promiscuous behaviour. Given the importance of education to enhance women empowerment and autonomy, it is plausible that such power dynamics among low educated WLWHA may negatively influence contraceptive use. Also, some barriers such as fear of side effects, safety of method, improper use and fear of violence by partner if disclosed have been reported in several studies and may explain the difference in contraceptive utilisation across educational levels. Another study in Ethiopia reported that with increasing education, women tend to value the benefits that contraception offers, which enhances acceptability.

Positive clustering tendencies were observed for levels of wealth index and education (see online supplementary file), which complemented the result obtained from the multiple logistic regression analysis. High clustering patterns with pockets of spatial outliers were seen for contraceptive non-use among study population in Nyanza and Nairobi. The high clustering pattern of contraceptive non-use seen in Nyanza and Nairobi might be explained by the densely populated geographical pattern while cluster outliers (HL and LH) in Nyanza, Western and Nairobi may be explained by epidemiological connectivity and demographic changes. Hotspot analysis in figure 6 further explains the reason for the spatial outliers seen in figure 3 by demonstrating the presence of high hotspots for low-income and middle-income study participants not using contraceptives aggregating in Nyanza and high-income hot spots with contraceptive non-use in Nairobi. In addition, there are low cluster/cold spot patterns seen in Rift Valley that may be due to more affluent study population living within those vicinities. Our findings validate other studies that reported low prevalence of contraceptive use in low-income WLWHA. Specifically, in a recent study conducted by Achana et al evaluating spatial patterns of sociodemographic determinants of contraceptive use in Ghanaian women, the authors reported higher levels of wealth to be predictive of contraceptive use. Harrington et al reported a low contraceptive uptake in Nyanza women accessing HIV care due to socioeconomic and demographic effects.

Cluster and hotspot analysis on levels of education with contraceptive non-use revealed strong pattern of clustering with rim of spatial outliers among study populations in Nyanza, Nairobi and Coast regions as seen in figure 4. Nyanza region had a high spatial clustering of contraceptive non-use among study participants with primary education. This finding supports previous research that examined spatial patterns and factors associated with increased need for family planning and its corresponding decreased uptake in Nigerian women, as determined from the 2013 Nigeria Demographic Health Survey (DHS). Moreover, Oyeronke and colleagues showed that women with higher educational level had high use of contraception. Likewise, another study that examined spatial patterns of contraceptive use in the Democratic Republic of Congo reported that individuals with no education were less likely to use contraceptives. Also, the map of WLWHA not using contraceptives (figure 3) follows the same pattern as figure 4. In other words, we were able to show with figures 3 and 4 that the study populations in regions of Kenya with low education had a higher concentration of non-contraceptive use. This important finding further elucidates the need for the incorporation of sexuality and reproductive health into school curriculum, as well as universal access to education for all women.
This conclusion is consistent with other studies that have posited that sexuality education can improve reproductive health outcomes among adolescents.13 14

Hotspots analysis in figure 6 supports the findings from the cluster analysis in figure 7 that Nairobi and Rift Valley are cold spots for low-income earners (low values surrounded by low values) of the study population not using contraceptives. These spatial patterns further reinforce the stark socioeconomic inequality and wealth index inherent within districts of Nairobi and Rift Valley when compared with poor localities in Nyanza. There were also spatial outliers (HL and LH) clusters of low-income population living in Nyanza and Western region of Kenya not using contraceptives. This can be explained in terms of socioeconomics, gender norms and inequalities regarding safe and transactional sex. Gaps in contraception utilisation are attributable to a divide in family planning programmes and HIV care services.8 36 Low contraceptive utilisation in WLWHA necessitates urgent action to prevent series of health consequences; therefore, a more concerted effort between health service providers and public health professionals in ensuring that PMTCT programmes focus on HIV antiretroviral treatment and on provision, and counselling on contraceptive usage, particularly on the dual method (the use of a barrier method such as condom with an additional modern contraception),38 is essential, enabling these women to make an informed decision.8 38–40

The results of this analysis underscore the significant role of sociodemographic factors as determinants to contraception use and subsequent prevention of adverse health events when these factors are addressed and provides target areas where resources should be directed to improve population health. Targeted interventions, such as educational and health awareness campaigns should place more emphasis on women who are socioeconomically disadvantaged. Advancing women education and empowerment is a key strategy to address misconceptions about contraception. The provision of counselling services on contraceptive use, HIV and family planning programme integration are possible measures that could be adopted to increase contraceptive usage among WLWHA. Also, efforts towards attaining substantial increase in contraceptive use in WLWHA must ensure long-term commitment to improve sustainability, and stringent rules should be kept in place for adequate accountability of funding and logistics of reproductive health programmes especially in African countries. An increase in contraceptive use mediated by adequate and accessible contraception provision is required to protect the reproductive and health needs of these women and cater for their unmet needs, in line with sections (iii) and (iv) of SDG of United Nations.41

The DHS data used in this study represent nationally representative data. Despite this strength, the study design was cross-sectional, and therefore the direction of causality cannot be ascertained. Furthermore, data collected were based on self-report and could be prone to recall and social desirability bias. However, given similarities of our results with other published studies, we postulate that potential biases had limited impact on the study findings. In the spatial analysis, the evidence of clustering was most observed for regions with higher sample size; there was minimal variation in the distribution of the sample population with sparse concentration in the Eastern and North Eastern part of the country when compared with other regions. Therefore, the local Moran’s Index used in our analysis could have overestimated or underestimated the effect of clustering tendencies for sociodemographic factors associated with contraceptive use. Similarly, this concern could have been responsible, at least in part, for the large chunk of insignificant clusters observed across some regions. Likewise, information and analysis derived from the DHS HIV-related dataset did not permit ascertainment of the potential epidemiological drivers of the significant spatial outliers observed in our spatial model.

Although the 2008–2009 DHS data used in this analysis are nearly a decade old, these data appear to represent the only publicly available data with HIV testing and GPS information for Kenya. Importantly, our results are consistent with conclusions from the Kenya AIDS Indicator Survey (KAIS) 2012 report,42 and results from this study based on currently obtainable data are presumed to still be largely reflective of current conditions, underscoring the urgent need of targeted intervention in hotspot areas.43 Numerous factors operating at various domains determine contraceptive use; however, our study examined the cluster patterns of these factors at the individual level. Future studies may focus on delineating the cluster patterns of structural and community level factors affecting contraceptive use among WLWHA.

CONCLUSION
These findings highlight the importance of extensive health education, counselling on safety of use of contraception, and improved access to contraception to mitigate the gap from social inequalities, including creation and enforcement of policies that protect and preserve women’s health in sub-Saharan African countries, especially Kenya, which are pertinent in achieving the SDG of promoting accessible and affordable reproductive services, gender equality and women empowerment.

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