Almost Optimal Distribution-free Junta Testing

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Abstract

We consider the problem of testing whether an unknown \( n \)-variable Boolean function is a \( k \)-junta in the distribution-free property testing model, where the distance between functions is measured with respect to an arbitrary and unknown probability distribution over \( \{0,1\}^n \). Chen, Liu, Servedio, Sheng and Xie [35] showed that the distribution-free \( k \)-junta testing can be performed, with one-sided error, by an adaptive algorithm that makes \( \tilde{O}(k^2/\epsilon) \) queries. In this paper, we give a simple two-sided error adaptive algorithm that makes \( \tilde{O}(k/\epsilon) \) queries.

1 Introduction

Property testing of Boolean function was first considered in the seminal works of Blum, Luby and Rubinfeld [11] and Rubinfeld and Sudan [42] and has recently become a very active research area. See for example, [1, 2, 3, 4, 7, 8, 13, 14, 15, 16, 18, 19, 22, 24, 27, 29, 32, 33, 37, 36, 39, 43] and other works referenced in the surveys [26, 40, 41].

A function \( f : \{0,1\}^n \to \{0,1\} \) is said to be \( k \)-junta if it depends on at most \( k \) variables. Juntas have been of particular interest to the computational learning theory community [9, 10, 12, 30, 34, 38]. A problem closely related to learning juntas is the problem of testing juntas: Given black-box query access to a Boolean function \( f \). Distinguish, with high probability, the case that \( f \) is \( k \)-junta versus the case that \( f \) is \( \epsilon \)-far from every \( k \)-junta.

In the uniform distribution framework, where the distance between two functions is measured with respect to the uniform distribution, Ficher et al. [24] introduced the junta testing problem and gave adaptive and non-adaptive algorithms that make \( \text{poly}(k)/\epsilon \) queries. Blais in [5] gave a non-adaptive algorithm that makes \( \tilde{O}(k^{3/2})/\epsilon \) queries and in [6] an adaptive algorithm that makes \( O(k \log k + k/\epsilon) \) queries. On the lower bounds side, Fisher et al. [24] gave an \( \Omega(\sqrt{k}) \) lower bound. Chockler and Gutfreund [21] gave an \( \Omega(k) \) lower bound for adaptive testing and, recently, Sağlam in [43] improved this lower bound to \( \Omega(k \log k) \). For the non-adaptive testing Chen et al. [17] gave the lower bound \( \Omega(k^{3/2})/\epsilon \).

In the distribution-free property testing, [28], the distance between Boolean functions is measured with respect to an arbitrary and unknown distribution \( D \) over \( \{0,1\}^n \). In this model, the testing algorithm is allowed (in addition to making black-box queries) to draw random \( x \in \{0,1\}^n \) according to the distribution \( D \). This model is studied in [20, 23, 25, 31, 35]. For testing \( k \)-junta in this model, Chen et al. [35] gave a one-sided adaptive algorithm that makes \( \tilde{O}(k^2)/\epsilon \) queries and proved a lower bound \( \Omega(2^{k/3}) \) for any non-adaptive algorithm. The results of Halevy and Kushilevitz [31] gives a one-sided non-adaptive algorithm that makes \( O(2^k/\epsilon) \) queries. The adaptive \( \Omega(k \log k) \) uniform-distribution lower bound from [43] trivially extend to the distribution-free model.

In this paper, we close the gap between the adaptive lower and upper bound. We prove

**Theorem 1.** For any \( \epsilon > 0 \), there is a two-sided distribution-free adaptive algorithm for \( \epsilon \)-testing \( k \)-junta that makes \( \tilde{O}(k/\epsilon) \) queries.
Our exact upper bound is $O((k/\varepsilon) \log(k/\varepsilon))$ and therefore, by Sağlam lower bound, $\Omega(k \log k)$ in [43], our bound is tight for any constant $\varepsilon$.

2 Preliminaries

In this section we give some notations follows by a formal definition of the model and some preliminary known results.

2.1 Notations

We start with some notations. Denote $[n] = \{1, 2, \ldots, n\}$. For $S \subseteq [n]$ and $x = (x_1, \ldots, x_n)$ we write $x(S) = \{x_i | i \in S\}$. For $X \subseteq [n]$ we denote by $\{0, 1\}^X$ the set of all binary strings of length $|X|$ with coordinates indexed by $i \in X$. For $x \in \{0, 1\}^n$ and $X \subseteq [n]$ we write $x_X \in \{0, 1\}^X$ to denote the projection of $x$ over coordinates in $X$. We denote by $1_X$ and $0_X$ the all one and all zero strings in $\{0, 1\}^X$, respectively. When we write $x_i = 0$ we mean $x_i = \bar{x}_i$. For $X_1, X_2 \subseteq [n]$ where $X_1 \cap X_2 = \emptyset$ and $x \in \{0, 1\}^{X_1} \cap \bar{y} \in \{0, 1\}^{X_2}$ we write $x \circ y$ to denote their concatenation, the string in $\{0, 1\}^{X_1 \cup X_2}$ that agrees with $x$ over coordinates in $X_1$ and agrees with $y$ over $X_2$. For $X \subseteq [n]$ we denote $\overline{X} = [n] \setminus X$. We say that the Boolean function $f : \{0, 1\}^n \rightarrow \{0, 1\}$ is a literal if $f \in \{x_1, \ldots, x_n, \bar{x}_1, \ldots, \bar{x}_n\}$.

Given $f, g : \{0, 1\}^n \rightarrow \{0, 1\}$ and a probability distribution $D$ over $\{0, 1\}^n$, we say that $f$ is $\varepsilon$-close to $g$ with respect to $D$ if $\Pr_{x \in D}[f(x) \neq g(x)] \leq \varepsilon$, where $x \in D$ means $x$ is chosen from $\{0, 1\}^n$ according to the distribution $D$. We say that $f$ is $\varepsilon$-far from $g$ with respect to $D$ if $\Pr_{x \in D}[f(x) \neq g(x)] \geq \varepsilon$. We say that $f$ is $\varepsilon$-far from every $k$-junta with respect to $D$ if for every $k$-junta $g$, $f$ is $\varepsilon$-far from $g$ with respect to $D$. We will use $U$ to denote the uniform distribution over $\{0, 1\}^n$.

2.2 The Model

In this subsection, we define the model.

We consider the problem of testing juntas in the distribution-free testing model. In this model, the algorithm has access to a $k$-junta $f$ via a black-box that returns $f(x)$ when a string $x$ is queried, and access to unknown distribution $D$ via an oracle that returns $x \in \{0, 1\}^n$ chosen randomly according to the distribution $D$.

A distribution-free testing algorithm $A$ is a algorithm that, given as input a distance parameter $\varepsilon$ and the above two oracles,

1. if $f$ is $k$-junta then $A$ output “accept” with probability at least $2/3$.
2. if $f$ is $\varepsilon$-far from every $k$-junta with respect to the distribution $D$ then it output “reject” with probability at least $2/3$.

We say that $A$ is one-sided if it always accepts when $f$ is $k$-junta, otherwise, it is called two sided algorithm. The query complexity of a distribution-free testing algorithm is the number of queries made on $f$.

2.3 Preliminaries Results

In this section, we give some known results that will be used in the sequel.

For a Boolean function $f$ and $X \subseteq [n]$, we say that $X$ is a relevant set of $f$ if there are $a, b \in \{0, 1\}^n$ such that $f(a) \neq f(b_X \circ a)$. When $X = \{i\}$ then we say that $x_i$ is relevant variable of $f$. Obviously, if $X$ is relevant set of $f$ then $x(X)$ contains at least one relevant variable of $f$. In particular, we have

Lemma 2. If $\{X_i\}_{i \in [r]}$ is a partition of $[n]$ then for any Boolean function $f$ the number of relevant sets $X_i$ of $f$ is at most the number of relevant variables of $f$. 

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We will use the following folklore result that is formally proved in [35].

**Lemma 3.** Let $\{X_i\}_{i \in [r]}$ be a partition of $[n]$. Let $f$ be a Boolean function and $u \in \{0,1\}^n$. If $f(u) \neq f(0)$ then a relevant set $X_\ell$ of $f$ with a string $v \in \{0,1\}^n$ that satisfies $f(v) \neq f(0_{X_\ell} \circ v_{X_\ell}^-)$ can be found with $\lceil \log_2 r \rceil$ queries.

The following is from [6].

**Lemma 4.** There exists a one-sided adaptive algorithm, UniformJunta$(f,k,\epsilon,\delta)$, for $\epsilon$-testing $k$-junta that makes $O(((k/\epsilon) + \log k) \log(1/\delta))$ queries and rejects $f$ with probability at least $1 - \delta$ when it is $\epsilon$-far from every $k$-junta with respect to the uniform distribution.

The following is from [35].

**Lemma 5.** Let $D$ be any probability distribution over $\{0,1\}^n$. If $f$ is $\epsilon$-far from every $k$-junta with respect to $D$ then for any $J \subseteq [n]$, $|J| \leq k$ we have

$$\Pr_{x \in D, y \in U}[f(x) \neq f(x \circ y)] \geq \epsilon.$$ 

**Proof.** Let $J \subseteq [n]$ of size $|J| \leq k$. For every fixed $y \in \{0,1\}^n$ the function $f(x \circ y)$ is $k$-junta and therefore

$$\Pr_{x \in D}[f(x) \neq f(x \circ y)] \geq \epsilon.$$ 

Therefore

$$\Pr_{x \in D, y \in U}[f(x) \neq f(x \circ y)] \geq \epsilon.$$ 

\[ \square \]

### 3 The Algorithm

In this section, we prove the correctness of the algorithm and show that it makes $\tilde{O}(k/\epsilon)$ queries. We first give an overview of the algorithm then prove its correctness and analyze its query complexity.

#### 3.1 Overview of the Algorithm

Consider the algorithm in Figure 1. In steps 1-2, the algorithm uniformly at random partitions $[n]$ into $r = O(k^2)$ disjoint sets $X_1, \ldots, X_r$. Lemma 6 shows that,

**Fact 1.** If the function is $k$-junta then with high probability (w.h.p), each set of variables $x(X_i) = \{x_j | j \in X_i\}$ contains at most one relevant variable.

In steps 3-12 the algorithm finds

**Fact 2.** relevant sets $\{X_i\}_{i \in I}$ such that for $X = \cup_{i \in I} X_i$, w.h.p., the function $f(x_X \circ 0_X)$ is $\epsilon/2$-close to $f$ with respect to $D$.

This is proved in Lemma 10. In addition, for each relevant set $X_\ell$, $\ell \in I$, it finds a string $v^{(\ell)}$ that satisfies $f(v^{(\ell)}) \neq f(0_{X_\ell} \circ v^{(\ell)}_{X_\ell})$. Obviously, if $|I| > k$ then, since each relevant set contains at least one relevant variable, the target is not $k$-junta and the algorithm rejects. See Lemma 2. Now one of the key ideas is the following: If $f$ is $k$-junta then $f(x_X \circ 0_X)$ is $k$-junta. If $f$ is $\epsilon$-far from every $k$-junta with respect to $D$ then since, by Fact 2 w.h.p., $f(x_X \circ 0_X)$ is $\epsilon/2$-close to $f$ with respect to $D$ we have that,

**Fact 3.** If $f$ is $\epsilon$-far from every $k$-junta with respect to $D$ then, w.h.p., $f(x_X \circ 0_X)$ is $\epsilon/2$-far from every $k$-junta with respect to $D$.

Since each $X_\ell$, $\ell \in I$ is relevant set, for $\ell \in I$ the function $f(x_{X_\ell} \circ v^{(\ell)}_{X_\ell})$ is non-constant. In steps 13-17 the algorithm tests that,
Fact 4. w.h.p., for each \( \ell \in I \) there is \( \tau(\ell) \in X_\ell \) such that \( f(x_{\tau(\ell)} \circ \tau(\ell)) \) is close to some literal in \( \{x_{\tau(\ell)}, \overline{x}_{\tau(\ell)}\} \), with respect to the uniform distribution.

This is done using the procedure \textbf{UniformJunta} in Lemma 4.

If \( f \) is \( k \)-junta then, by Fact 1 w.h.p., it passes this test (does not output reject). This is Lemma 7. If the algorithm does not pass this test, it rejects. If \( f \) is not \( k \)-junta and it passes this test, then the statement in Fact 4 is true. This is proved in Lemma 11.

Consider now steps 18-28. First, let us consider a function \( f \) that is \( \epsilon \)-far from every \( k \)-junta with respect to \( D \). Let \( J = \{\tau(\ell) \mid \ell \in I\} \) where \( \tau(\ell) \) is as defined in Fact 3. Since by Fact 3 w.h.p., \( f(x_j \circ \overline{0}_X) \) is \( \epsilon \)-far from every \( k \)-junta with respect to \( D \) and \( |J| = |I| \leq k \), by Lemma 5 w.h.p.,

\[
\Pr_{y \in U, x \in D}[f(x_j \circ \overline{0}_{X}) \neq f(x_j \circ y_{X \setminus J} \circ \overline{0}_X)] \geq \epsilon/2.
\]

This is the last test we would like to do but the problem is that we do not know \( J \), so we cannot use this test as is. We change it, as is done in [32], to an equivalent test as follows

\[
\Pr_{z \in U, x \in D}[f(x_j \circ \overline{0}_X) \neq f((x_j + x) \circ \overline{0}_X) \mid z_J = 0, j] \geq \epsilon/2.
\]

To be able to draw uniformly random \( x_X \) with \( z_J = 0 \), we use Fact 4 that is, the fact that each \( f(x_{\tau(\ell)} \circ \tau(\ell)) \) is close to one of the literals in \( \{x_{\tau(\ell)}, \overline{x}_{\tau(\ell)}\} \). For every \( \ell \in I \) the algorithm draws uniformly random \( w := x_{\tau(\ell)} \)
and then using the fact that \( f(x_{\tau(\ell)} \circ \tau(\ell)) \) is close to one of the literals in \( \{x_{\tau(\ell)}, \overline{x}_{\tau(\ell)}\} \) where \( \tau(\ell) \in X_\ell \) the algorithm tests in which set \( Y_{\ell,0} := \{j \in X_\ell \mid w_j = 0\} \) or \( Y_{\ell,1} := \{j \in X_\ell \mid w_j = 1\} \) the index \( \tau(\ell) \) falls. If \( \tau(\ell) \in Y_{\ell,0} \) then the entry \( \tau(\ell) \) in \( x_{\tau(\ell)} \) is zero and if \( \tau(\ell) \in Y_{\ell,1} \) then the entry \( \tau(\ell) \) in \( x_{\tau(\ell)} \) is one. In the latter case, the algorithm replaces \( x_{\tau(\ell)} \) with \( \overline{x}_{\tau(\ell)} \). This gives a random uniform \( x_{\tau(\ell)} \) with \( x_{\tau(\ell)} = 0 \). We do that for every \( \ell \in I \) and get a random uniform \( x \) with \( z_J = 0 \). This is proved in Lemma 12. Then the algorithm rejects if \( f(x_X \circ \overline{0}_X) \neq f((x_X + x) \circ \overline{0}_X) \). If \( f(x_X \circ \overline{0}_X) \) is \( \epsilon/2 \)-far from every \( k \)-junta, the algorithm reject with probability at least \( \epsilon/2 \). Therefore, by repeating this test \( O(1/\epsilon) \) times the algorithm rejects w.h.p. This is proved in Lemma 13.

Now we consider \( f \) that is \( k \)-junta. Obviously, if \( f \) is \( k \)-junta then \( f(x_X \circ \overline{0}_X) = f((x_X + x_X) \circ \overline{0}_X) \) when \( z_J = 0 \) and the algorithm accepts. This is because \( x(J) \) are the relevant variables in \( f(x_X \circ \overline{0}_X) \). This is proved in Lemma 8.

3.2 The algorithm for \( k \)-Junta

In this subsection, we show that if the target function \( f \) is \( k \)-junta then the algorithm accepts with probability at least \( 2/3 \).

We first prove

Lemma 6. Consider steps 13 in the algorithm. If \( f \) is a \( k \)-junta then, with probability at least \( 2/3 \), for each \( i \in [r] \), the set \( x(X_i) = \{x_j \mid j \in X_i\} \) contains at most one relevant variable of \( f \).

Proof. Let \( x_{i_1} \) and \( x_{i_2} \) be two relevant variables in \( f \). The probability that \( x_{i_1} \) and \( x_{i_2} \) are in the same set is equal to \( 1/r \). By the union bound, it follows that the probability that some relevant variables \( x_{i_1} \) and \( x_{i_2} \) in \( f \) are in the same set is at most \( (k/5)/r \leq 1/3 \). \( \square \)

We now show that w.h.p. the algorithm reaches the final test in the algorithm

Lemma 7. If \( f \) is \( k \)-junta and each \( x(X_i) \) contains at most one relevant variable of \( f \) then

1. Each \( x(X_i), i \in I \), contains exactly one relevant variable.

2. The algorithm reaches step 18.
Algorithm Simple\textsubscript{Dk–Junta}$(f, D, \epsilon)$

\textit{Input}: Oracle that access a Boolean function $f$ and
\textit{oracle} that chooses $x \in \{0, 1\}^n$ according to the distribution $D$.

\textit{Output}: Either “accept” or “reject”

Partition $[n]$ into $r$ sets
1. Set $r = 2k^2$.
2. Choose uniformly at random a partition $X_1, X_2, \ldots, X_r$ of $[n]$

Find a close function and relevant sets
3. Set $X = \emptyset; I = \emptyset; t(X) = 0$
4. Repeat $M = 2k \ln(15k)/\epsilon$ times
5. Choose $u \in D$.
6. $t(X) \leftarrow t(X) + 1$
7. If $f(u_X \circ 0_{X}) \neq f(u)$ then
   \hspace{1em} Find a new relevant set $X_\ell$; $X \leftarrow X \cup X_\ell$; $I \leftarrow I \cup \{\ell\}$
8. \hspace{1em} Find a string $v(\ell) \in \{0, 1\}^n$ such that $f(v(\ell)) \neq f(0_{X_\ell} \circ 0_{X})$
9. \hspace{1em} If $|I| > k$ then output “reject” and halt.
10. \hspace{1em} $t(X) = 0$.
11. \hspace{1em} If $t(X) = 2 \ln(15k)/\epsilon$ then Goto 13.

Tests if each relevant set corresponds to a Boolean function that is close to a literal
13. For every $\ell \in I$ do
14. \hspace{1em} If Uniform\textsubscript{Junta}$f(x_{X_\ell} \circ 0_{X}) = 1/30, 1/15)$=“reject”
15. \hspace{1em} then output “reject” and halt
16. \hspace{1em} Choose $b \in U$
17. \hspace{1em} If $f(b_{X_\ell} \circ 0_{X}) = f(b_{X_\ell} \circ 0_{X})$ then output “reject” and halt

The final test of Lemma 5
18. Repeat $M' = (2 \ln 15)/\epsilon$ times
19. \hspace{1em} Choose $w \in U$; $z = 0_{X}$
20. \hspace{1em} For every $\ell \in I$ do
21. \hspace{2em} Set $Y_{\ell, \xi} = \{ j \in X_\ell | w_j = \xi \}$ for $\xi \in \{0, 1\}$.
22. \hspace{2em} Set $G_{\ell, 0} = G_{\ell, 1} = 0$;
23. \hspace{2em} Repeat $h = \ln(15M'k)/\ln(4/3)$ times
24. \hspace{3em} Choose $b \in U$;
25. \hspace{4em} If $f(b_{Y_{\ell, 0}} \circ b_{Y_{\ell, 1}}) \neq f(b_{Y_{\ell, 0}} \circ b_{Y_{\ell, 1}})$ then $G_{\ell, 0} \leftarrow G_{\ell, 0} + 1$
26. \hspace{4em} If $f(b_{Y_{\ell, 1}} \circ b_{Y_{\ell, 0}}) \neq f(b_{Y_{\ell, 1}} \circ b_{Y_{\ell, 0}})$ then $G_{\ell, 1} \leftarrow G_{\ell, 1} + 1$
27. \hspace{2em} If $( \{ G_{\ell, 0}, G_{\ell, 1} \} \neq \{0, h\})$ then output “reject” and halt.
28. \hspace{2em} Choose $u \in D$
29. \hspace{2em} If $f(u_X \circ 0_{X}) \neq f((u_X + z_X) \circ 0_{X})$ then output “reject” and halt.
30. Output “accept”

Figure 1: A two-sided distribution-free adaptive algorithm for $\epsilon$-testing $k$-junta.

Proof. By Lemma 8 and steps 7\textsuperscript{9} for $\ell \in I, f(v(\ell)) \neq f(0_{X_\ell} \circ 0_{X})$ and $x(X_\ell)$ contains exactly one relevant
variable. Therefore, for every \( \ell \in I \), \( f(x_{X_\ell} \circ v_{\ell}(t)) \) is a literal.

If the algorithm does not reach step 18 then it either halts in step 10 or 17. If it halts in step 10 then \(|I| > k\) and therefore, by Lemma 2, \( f \) contains more than \( k \) relevant variables and then it is not \( k \)-Junta. If it halts in step 17 then, by Lemma 11 for some \( X_{\ell} \), \( \ell \in I \), \( f(x_{X_\ell} \circ v_{\ell}(t)) \) is not 1-Junta (literal or constant function) and therefore \( X_{\ell} \) contains at least two relevant variables. If it halts in step 17 then \( f(b_{X_\ell} \circ v_{\ell}(t)) = f(b_{X_\ell} \circ v_{\ell}(t)) \) and then \( f(x_{X_\ell} \circ v_{\ell}(t)) \) is not a literal. In all cases we get a contradiction.

We now give two Lemmas that show that, with probability at least \( 2/3 \), the algorithm accepts.

Lemma 8. If \( f \) is \( k \)-Junta and each \( x(X_\ell) \) contains at most one relevant variable of \( f \) then the algorithm outputs “accept”.

Proof. By Lemma 7 the algorithm reaches step 18. We now show that it reaches step 29. Now we need to show that the algorithm does not halt in step 25 or 28.

Since \( Y_{\ell,0}, Y_{\ell,1} \) is a partition of \( X_{\ell} \), \( \ell \in I \) and \( X_{\ell} \) contains exactly one relevant variable in \( x(X_\ell) \) of \( f \), this variable is either in \( x(Y_{\ell,0}) \) or in \( x(Y_{\ell,1}) \) but not in both. Suppose w.l.o.g. it is in \( x(Y_{\ell,0}) \) and not in \( x(Y_{\ell,1}) \).

Then \( f(x_{Y_{\ell,0}} \circ b_{Y_{\ell,0}} \circ v_{\ell}(t)) \) is a literal and \( f(x_{Y_{\ell,1}} \circ b_{Y_{\ell,0}} \circ v_{\ell}(t)) \) is a constant function. This implies that for any \( b \), \( f(b_{Y_{\ell,0}} \circ b_{Y_{\ell,1}} \circ v_{\ell}(t)) \neq f(b_{Y_{\ell,0}} \circ b_{Y_{\ell,1}} \circ v_{\ell}(t)) \) and \( f(b_{Y_{\ell,0}} \circ b_{Y_{\ell,1}} \circ v_{\ell}(t)) = f(b_{Y_{\ell,0}} \circ b_{Y_{\ell,0}} \circ v_{\ell}(t)) \). Therefore, \( G_{\ell,0} = h \) and \( G_{\ell,1} = 0 \). Thus the algorithm does not halt in step 25.

Now for every \( X_{\ell} \), \( \ell \in I \), let \( \tau(t) \in X_{\ell} \) be such that \( f(x_{X_\ell} \circ v_{\ell}(t)) \in \{x_{\tau(t)}, x_{\tau(t)}\} \). If \( \tau(t) \in Y_{\ell,0} \) then \( G_{\ell,0} = h \) and then by step 29, \( z_{\tau(t)} = u_{\tau(t)} = 0 \). If \( \tau(t) \in Y_{\ell,1} \) then \( G_{\ell,1} = h \) and then \( z_{\tau(t)} = u_{\tau(t)} = 0 \). Therefore for every relevant variable \( x_{\tau(t)} \) in \( f \) we have \( z_{\tau(t)} = 0 \) which implies that \( f(u_X \circ 0_{X'}) = f((u_X + z_X) \circ 0_{X'}) \) and therefore the algorithm does not halt in step 28.

Lemma 9. If \( f \) is \( k \)-Junta then the algorithm outputs “accept” with probability at least \( 2/3 \).

Proof. The result follows from Lemma 6 and Lemma 8.

3.3 The Algorithm for \( \epsilon \)-Far Functions

In this subsection, we prove that if \( f \) is \( \epsilon \)-far from every \( k \)-junta then the algorithm rejects with probability at least \( 2/3 \).

The first lemma shows that, w.h.p., \( f(u_X \circ 0_{X'}) \) is \( \epsilon/2 \)-close to \( f \).

Lemma 10. If the algorithm reaches step 13 then \( t(X) = 2 \ln(15k)/\epsilon \) and \(|I| \leq k \). If

\[
\Pr_{u \in \mathcal{D}}[f(u_X \circ 0_{X'}) \neq f(u)] \geq \epsilon/2
\]

then, with probability at most \( 1/15 \), the algorithm reaches step 13.

Proof. The algorithm does not reach step 13 if and only if it halts in step 10 and then \(|I| > k \). The size of \( I \) is increased by one each time the condition, \( f(u_X \circ 0_{X'}) \neq f(u) \), in step 7 is true. Therefore, if the algorithm reaches step 13 then the condition in step 4 was true at most \( k \) times and \(|I| \leq k \). Then steps 8-14 are executed at most \( k \) times. Thus, \( t() \) is updated to 0 at most \( k \) times. The loop 4-12 is repeated \( M \) times and \( t() \) is updated to 0 at most \( k \) times and therefore there is \( X \) for which \( t(X) = M/k = 2 \ln(15k)/\epsilon \). This implies that when the algorithm reaches step 13 we have \( t(X) = 2 \ln(15k)/\epsilon \).

The probability that the algorithm reaches step 13 with \( \Pr_{u \in \mathcal{D}}[f(u_X \circ 0_{X'}) \neq f(u)] > \epsilon/2 \) is the probability that for one (of the at most \( k \)) \( X' \), \( \Pr_{u \in \mathcal{D}}[f(u_X \circ 0_{X'}) \neq f(u)] > \epsilon/2 \) and \( t(X') = 2 \ln(15k)/\epsilon \). By the union bound, this probability is less than

\[
k \left(1 - \frac{\epsilon}{2}\right)^{2 \ln(15k)/\epsilon} = \frac{1}{15}\]

\( \square \)
In the following lemma we show that, w.h.p, each \( f(x_{X_t} \circ v^{(f)}_{X_t}) \) is close to a literal.

**Lemma 11.** Consider steps 15, 16 If for some \( \ell \in I \), \( f(x_{X_t} \circ v^{(f)}_{X_t}) \) is \((1/30)\)-far from every literal with respect to the uniform distribution then, with probability at least \( 1 - (2/15) \), the algorithm reject.

**Proof.** If \( f(x_{X_t} \circ v^{(f)}_{X_t}) \) is \((1/30)\)-far from every literal with respect to the uniform distribution then it is either (case 1) \((1/30)\)-far from every 1-Junta (literal or constant) or (case 2) \((1/30)\)-far from every literal and \((1/30)\)-close to 0-Junta. In case 1, by Lemma 3, with probability at least \( 1 - (1/15) \), UniformJunta \( (f(x_{X_t} \circ v^{(f)}_{X_t}), 1, 1/30, 1/15) \) = “reject” and the then algorithm rejects. In case 2, if \( f(x_{X_t} \circ v^{(f)}_{X_t}) \) is \((1/30)\)-close to some 0-Junta then it is either \((1/30)\)-close to 0 or \((1/30)\)-close to 1. Suppose it is \((1/30)\)-close to 0. Let \( b \) be a random uniform string generated in steps 16. Then \( \overline{b} \) is random uniform and for \( g(x) = f(x_{X_t} \circ v^{(f)}_{X_t}) \) we have

\[
\Pr[\text{The algorithm does not reject}] = \Pr[g(b) \neq g(\overline{b})] = \Pr[g(b) = 1 \land g(\overline{b}) = 0] + \Pr[g(b) = 0 \land g(\overline{b}) = 1] \\
\leq \Pr[g(b) = 1] + \Pr[g(\overline{b}) = 1] \\
\leq \frac{1}{15}.
\]

By the union bound the result follows. \( \square \)

In the next lemma we prove that, w.h.p, the string \( z \) generated in steps 19–20 satisfies \( z_J = 0 \) where \( x(J) \) are relevant variables of \( f(u_X \circ 0_X) \).

**Lemma 12.** Consider steps 19, 20 If for every \( \ell \in I \) the function \( f(x_{X_t} \circ v^{(f)}_{X_t}) \) is \((1/30)\)-close to a literal in \( \{x_{\tau(\ell)}, \overline{x}_{\tau(\ell)}\} \) with respect to the uniform distribution, where \( \tau(\ell) \in X_\ell \), and \( \{G_{\ell,0}, G_{\ell,1}\} = \{0, h\} \) then, with probability at least \( 1 - k(3/4)^h \), we have: For every \( \ell \in I \), \( z_{\tau(\ell)} = 0 \).

**Proof.** Fix some \( \ell \). Suppose \( f(x_{X_t} \circ v^{(f)}_{X_t}) \) is \((1/30)\)-close to \( x_{\tau(\ell)} \) with respect to the uniform distribution. The case when it is \((1/30)\)-close to \( \overline{x_{\tau(\ell)}} \) is similar. Since \( X_\ell = Y_{\ell,0} \cup Y_{\ell,1} \) and \( Y_{\ell,0} \cap Y_{\ell,1} = \emptyset \) we have that \( \tau(\ell) \in Y_{\ell,0} \) or \( \tau(\ell) \in Y_{\ell,1} \), but not both. Suppose \( \tau(\ell) \in Y_{\ell,0} \). The case where \( \tau(\ell) \in Y_{\ell,1} \) is similar. Define the random variable \( Z(x_{X_t}) = 1 \) if \( f(x_{X_t} \circ v^{(f)}_{X_t}) \neq x_{\tau(\ell)} \) and \( Z(x_{X_t}) = 0 \) otherwise. Then

\[
\mathbf{E}_{x_{X_t} \in U} [Z(x_{X_t})] \leq \frac{1}{30}.
\]

Therefore

\[
\mathbf{E}_{x_{Y_{\ell,1}} \in U} \mathbf{E}_{x_{Y_{\ell,0}} \in U} [Z(x_{Y_{\ell,0}} \circ x_{Y_{\ell,1}})] \leq \frac{1}{30}
\]

and by Markov’s bound

\[
\Pr_{x_{Y_{\ell,1}} \in U} \left[ \mathbf{E}_{x_{Y_{\ell,0}} \in U} [Z(x_{Y_{\ell,0}} \circ x_{Y_{\ell,1}})] \geq \frac{2}{15} \right] \leq \frac{1}{4}.
\]

That is, for a random uniform string \( b \in \{0, 1\}^n \), with probability at least \( 3/4 \), \( f(x_{Y_{\ell,0}} \circ b_{Y_{\ell,1}} \circ v^{(f)}_{X_t}) \) is \((2/15)\)-close to \( x_{\tau(\ell)} \) with respect to the uniform distribution. Now, given that \( f(x_{Y_{\ell,0}} \circ b_{Y_{\ell,1}} \circ v^{(f)}_{X_t}) \) is \((2/15)\)-close to \( x_{\tau(\ell)} \) with respect to the uniform distribution the probability that \( G_{\ell,0} = 0 \) is the probability that \( f(b_{Y_{\ell,0}} \circ b_{Y_{\ell,1}} \circ v^{(f)}_{X_t}) = f(b_{Y_{\ell,0}} \circ b_{Y_{\ell,1}} \circ v^{(f)}_{X_t}) \) for \( h \) random uniform strings \( b \in \{0, 1\}^n \). Let \( b^{(1)}, \ldots, b^{(h)} \) be \( h \) random uniform strings in \( \{0, 1\}^n \), \( V(b) \) be the event \( f(b_{Y_{\ell,0}} \circ b_{Y_{\ell,1}} \circ v^{(f)}_{X_t}) = f(b_{Y_{\ell,0}} \circ b_{Y_{\ell,1}} \circ v^{(f)}_{X_t}) \) and \( A \)
the event that \( f(x_{Y_{\ell,0}} \circ b_{Y_{\ell,1}} \circ i_{X_{\ell,0}}^{(i)}) \) is \((2/15)\)-close to \( x_{\tau(\ell)} \) with respect to the uniform distribution. Let 
\[ g(x_{Y_{\ell,0}}) = f(x_{Y_{\ell,0}} \circ b_{Y_{\ell,1}} \circ i_{X_{\ell,0}}^{(i)}). \]
Then
\[
\Pr[V(b)|A] = \Pr[g(b_{Y_{\ell,0}}) = g(b_{Y_{\ell,0}})|A] \\
= \Pr[(g(b_{Y_{\ell,0}}) = b_{\tau(\ell)} \vee g(b_{Y_{\ell,0}}) = b_{\tau(\ell)}) \vee (g(b_{Y_{\ell,0}}) = b_{\tau(\ell)} \wedge g(b_{Y_{\ell,0}}) = \overline{b_{\tau(\ell)}})|A] \\
\leq \Pr[g(b_{Y_{\ell,0}}) \neq b_{\tau(\ell)} \wedge g(b_{Y_{\ell,0}}) \neq \overline{b_{\tau(\ell)}}|A] \\
\leq \Pr[g(b_{Y_{\ell,0}}) \neq b_{\tau(\ell)}|A] + \Pr[g(b_{Y_{\ell,0}}) \neq \overline{b_{\tau(\ell)}}|A] \leq \frac{4}{15}.
\]
Since \( \tau(\ell) \in Y_{\ell,0} \), we have \( w_{\tau(\ell)} = 0 \). Therefore, by step 26 and since \( \tau(\ell) \in X_{\ell} \),
\[
\Pr[z_{\tau(\ell)} = 1] = \Pr[G_{\ell,0} = 0 \wedge G_{\ell,1} = \overline{h}] \\
\leq \Pr[G_{\ell,0} = 0] = \Pr[(\forall j \in [h]) V(b^{(j)})] \\
= (\Pr[V(b)])^h \leq (\Pr[V(b)|A] + \Pr[A])^h \leq (4/15 + 1/4)^h \leq (3/4)^h.
\]
Therefore, the probability that \( z_{\tau(\ell)} = 1 \) for some \( \ell \in I \) is at most \( k(3/4)^h \).

We now show that w.h.p the algorithm rejects.

**Lemma 13.** If \( f \) is \( \epsilon \)-far from every \( k \)-junta with respect to \( D \) then, with probability at least \( 2/3 \), the algorithm outputs “reject”.

**Proof.** If the algorithm stops in step 13 then we are done. Therefore we may assume that
\[ |I| \leq k. \]
By Lemma 11 if \( \Pr_{u \in D}[f(x \circ 0_{X}) \neq f(u)] \geq \epsilon/2 \) then, with probability at most \( 1/15 \), the algorithm reaches step \( \ell \). So we may assume that (failure probability \( 1/15 \))
\[ \Pr_{u \in D}[f(x \circ 0_{X}) \neq f(u)] \leq \epsilon/2. \]
Since \( f \) is \( \epsilon \)-far from every \( k \)-junta with respect to \( D \) and \( f(x \circ 0_{X}) \) is \( \epsilon/2 \)-close to \( f \) with respect to \( D \) we have \( f(x \circ 0_{X}) \) is \( \epsilon/2 \)-far from every \( k \)-junta with respect to \( D \). Therefore, by Lemma 5
\[ \Pr_{u \in D, y \in U}[f(x \circ 0_{X})] = f(u \circ y_{X \setminus I} \circ 0_{X}) \geq \frac{\epsilon}{2}. \]
By Lemma 11 if some \( f(x \circ i_{X}^{(i)}) \) is \( (1/30) \)-far from any literal with respect to the uniform distribution then, with probability at least \( 1 - (2/15) \), the algorithm rejects. So we may assume (failure probability \( 2/15 \)) that every \( f(x \circ i_{X}^{(i)}) \) is \( (1/30) \)-close to some \( x_{\tau(\ell)} \) or \( x_{\tau(\ell)} \) with respect to the uniform distribution, where \( \tau(\ell) \in X_{\ell} \).

Let \( z^{(1)}, \ldots, z^{(M')} \) be the strings generated in step 26. By Lemma 12, with probability at least \( 1 - M'k(3/4)^h \geq 1 - (1/15) \), every \( z^{(i)} \) generated in step 26 satisfies \( z^{(i)}_{\tau(\ell)} = 0 \) for all \( \ell \in I \). Also, since the distribution of \( w_{X_{\ell}} \) and \( \overline{w_{X_{\ell}}} \) is uniform, the distribution of \( z^{(i)}_{X \setminus I} \) and \( u_{X \setminus I} + z^{(i)}_{X \setminus I} \) is uniform. We now assume (failure probability \( 1/15 \)) that \( z^{(i)}_{I} = 0 \) for all \( i \). Therefore, by (3),
\[
\Pr_{u \in D, z^{(i)}_{X \setminus I} \in U}[(\forall i) f(u \circ 0_{X}) = f((u \circ z^{(i)}_{X}) \circ 0_{X})] \\
= \left( \Pr_{u \in D, z^{(i)}_{X \setminus I} \in U}[f(u \circ z^{(i)}_{X}) = f((u \circ z^{(i)}_{X}) \circ 0_{X})] \right)^{M'} \\
= \left( \Pr_{u \in D, y \in U}[f(u \circ y_{X \setminus I} \circ 0_{X})] \right)^{M'} \\
\leq (1 - \epsilon/2)^{M'} \leq \frac{1}{15}.
\]
Therefore, the failure probability of an output “reject” is at most $1/15 + 2/15 + 1/15 + 1/15 = 1/3$.

### 3.4 The Query Complexity of the Algorithm

In this section we show that

**Lemma 14.** The query complexity of the algorithm is

$$\tilde{O}\left(\frac{k}{\epsilon}\right).$$

*Proof.* The condition in step 4 requires two queries and is executed at most $M = 2k\ln(15k)/\epsilon$ times. This is $2M = O((k \log k)/\epsilon)$ queries. Steps 8 is executed at most $k + 1$ times. This is because each time it is executed, the value of $|I|$ is increased by one, and when $|I| = k + 1$ the algorithm rejects. By Lemma 3 to find a new relevant set the algorithm makes $O(\log r) = O(\log k)$ queries. This is $O(k \log k)$ queries. Steps 13 and 17 are executed $|I| \leq k$ times, and by Lemma 4 the total number of queries made is $O(1/(1/30) \log(15))k + 2k = O(k)$.

The final test in the algorithm is repeated $M' = (2 \ln 15)/\epsilon$ times (step 18) and each time, and for each $\ell \in I$, (step 20) it repeats $h$ times (step 23) two conditions that takes 2 queries each (step 24). This takes $4M'kh = O((k/\epsilon) \ln(k/\epsilon))$ queries. The number of queries in step 28 is $2M' = O(1/\epsilon)$. Therefore the total number of queries is

$$O\left(\frac{k}{\epsilon} \ln \frac{k}{\epsilon}\right).$$

\[\square\]

### 4 Open Problems

In this paper we proved that for any $\epsilon > 0$, there is a two-sided distribution-free adaptive algorithm for $\epsilon$-testing $k$-junta that makes $\tilde{O}(k/\epsilon)$ queries. It is also interesting to find a one-sided distribution-free adaptive algorithm with such query complexity.

Chen et al. [35] proved the lower bound $\Omega(2^{k/3})$ for any non-adaptive (one round) algorithm. What is the minimal number rounds one needs to get $\text{poly}(k/\epsilon)$ query complexity? Can $O(1)$-round algorithms solve the problem with $\text{poly}(k/\epsilon)$ queries?

In the uniform distribution framework, where the distance between two functions is measured with respect to the uniform distribution Blais in [5] gave a non-adaptive algorithm that makes $\tilde{O}(k^{3/2})/\epsilon$ queries and in [4] an adaptive algorithm that makes $O(k \log k + k/\epsilon)$ queries. On the lower bounds side, Chockler and Gutfreund [21] gave an $\Omega(k)$ lower bound for adaptive testing and Chen et al. [17] gave an $\tilde{\Omega}(k^{3/2})/\epsilon$ lower bound for the non-adaptive testing. Thus in both the adaptive and non-adaptive uniform distribution settings, the query complexity of $k$-junta testing has now been pinned down to within logarithmic factors. It is interesting to study $O(1)$-round algorithms. For example, what is the query complexity for 2-round algorithm.

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