Impact of missing individual patient data on 18 meta-analyses of randomised trials in oncology: Gustave Roussy experience

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ABSTRACT

Objective To compare the characteristics, quality and treatment effects of randomised clinical trials (RCTs) by individual patient data (IPD) availability, in trials eligible for 18 IPD meta-analyses (MA).

Design Trial characteristics, risk of bias (RoB) and hazard ratio (HR) for overall survival were extracted from IPD-MA publications and/or RCTs publications. Data for the RoB assessment were extracted for a subset of 73 RCTs. Two investigators blinded to whether IPD was available or not evaluated the RoB for these trials. Treatment effects were compared using ratios of global HRs (RHRs) of IPD-unavailable trials and IPD-available trials. RHR were pooled using a fixed-effect model.

Data sources We examined the IPD availability for each trial eligible for each IPD-MA; when the IPD was not available for a trial, we used information from published sources.

Eligibility criteria for selecting studies We selected all published IPD-MAs conducted at Gustave Roussy and the RCTs eligible for each.

Results 349 RCTs (73 018 patients) from 18 MAs were included in each RCT. In collaborative work, the trial data sets are typically provided to a coordinating group by trial investigators. IPD-MAs of RCTs have many advantages compared with aggregated data MAs. First, they limit selective reporting when they include unpublished RCTs, excluded patients and unpublished or misreported outcomes. Moreover, access by the MA coordinators to the IPD allows checking of data quality and integrity and correction and completion of the data in collaboration with investigators. Finally, by standardising coding across RCTs and accessing data on covariates, IPD-MAs allow investigators to study new research questions, such as potential interactions between patient characteristics and the treatment effect.

Conclusions IPD-unavailable RCTs were significantly different from IPD-available RCTs in terms of trial characteristics and were at greater RoB. IPD-unavailable RCTs had a significantly greater treatment effect.

INTRODUCTION

Randomised clinical trials (RCTs) and meta-analyses (MA) of RCTs based on systematic review provide the highest level of evidence for assessing intervention effectiveness. There are two types of MAs: aggregated data MA and individual patient data MAs (IPD-MAs). The first and most common type is based on aggregated data (ie, data from people enrolled in a trial have been statistically combined) extracted from public sources (eg, journal publications). IPD-MA is based on original data from each patient included in each RCT. In collaborative work, the IPD allows checking of data quality and integrity and correction and completion of the data in collaboration with investigators. Moreover, access by the MA coordinators to the IPD allows checking of data quality and integrity and correction and completion of the data in collaboration with investigators.

Conclusions IPD-unavailable RCTs were significantly different from IPD-available RCTs in terms of trial characteristics and were at greater RoB. IPD-unavailable RCTs had a significantly greater treatment effect.
have IPD available, which can introduce a selection bias if IPD-unavailable RCTs differ from IPD-available RCTs, either in treatment effect or factors possibly related to the treatment effect. To compensate for the lack of IPD, MA researchers may elect to extract aggregated data from the publications of IPD-unavailable trials, although we do not at Gustave Roussy. For substitution of IPD by aggregated data to be justifiable, however, meta-analysable data must be available in publications and the extracted aggregated data must be from unbiased trials. Then, their quality should be evaluable.

Our objective was to examine the IPD-MAs conducted by the Gustave Roussy Meta-analysis Unit, which has served as the coordinating group, to estimate the proportions of eligible RCTs in each MA that were IPD unavailable and IPD available, to compare the trial characteristics and risk of bias (RoB) (trial ‘quality’) of IPD-unavailable with IPD-available RCTs and to compare, across the MAs, summary treatment effects based on the pooling of published aggregated data from IPD-unavailable RCTs and summary treatment effects from the IPD-MA publications.

METHODS
Eligible meta-analysis and RCT selection
Gustave Roussy Meta-analysis Unit has coordinated and published 18 IPD-MAs based on RCTs since 1990; all aimed to estimate the effect of a treatment combination on overall survival in cancer patients (see online supplementary etable 1). Every MA we have published was eligible for our study. RCT selection for each MA was based on systematic review and has been previously described.

Standard procedures were used to check the RCT data sets sent to us for the IPD-MA, in collaboration with the investigators (online supplement). If we had major doubts about a trial methodological quality based on its protocol and our re-analyses of its IPD (see online supplementary etable 2) (eg, allocation concealment, missing data), we had the option of excluding it from the IPD-MA (‘excluded IPD-available RCTs’). For each MA we conducted, we described all eligible RCTs, regardless of whether they were IPD-available, IPD-unavailable or excluded IPD-available RCTs (see online supplementary etable 1). All the publications of 18 IPD-MAs included only IPD-available RCTs for treatment evaluation, however.

Data collection
For every eligible RCT, we collected the following data from the available publications: number of randomised patients, date of first patient randomised, date of publication, countries involved, type of publication (eg, conference abstract), nature of funding, whether the RCT was multicentre or not and whether it was international or not. This information was extracted from RCT publications for IPD-unavailable RCTs and from IPD-MA publications for RCTs with IPD supplemented by the RCT publications and/or our archives if necessary.

To assess the RoB in IPD-unavailable RCTs and compare it with the risk in IPD-available RCTs, we first selected RCTs reported in English as full-text articles and excluded conference abstracts and unpublished RCTs because methodological information was insufficient most of the time. Second, IPD-available RCT publication(s) were selected to match with IPD-unavailable RCT publication on MA to which they belonged and the period when the RCT was performed based on the date of randomisation of the first patient (before 1997/1997 and after). If more than two IPD-available RCTs by IPD-unavailable RCT were available in the strata defined by the two above criteria, we used the Excel ALEA function to randomly select publication of two IPD-available RCTs by IPD-unavailable RCT publication. The number of IPD-available RCT publications selected was limited because of time constraints. Two investigators blinded to whether IPD was available or not extracted data from the selected publications for both IPD-unavailable and the sample of IPD-available RCTs to evaluate RoB using the Cochrane tool. Five dimensions were studied and graded as low risk, high risk or unclear: random sequence generation, allocation concealment, blinding of personnel and participants, blinding of outcome assessment and incomplete outcome data. Disagreements were resolved by discussion with a third investigator.

For IPD-unavailable RCTs, we extracted the HR of treatment effect on overall survival and its 95% CI from each RCT publication. If the hazard ratio (HR) and its 95% confidence interval (CI) were not reported, we calculated them using one of two validated methods depending on available information. For IPD-available RCTs, we extracted the global HR on overall survival and its 95% CI from the MA publication, using the most recent when there was an updated IPD-MA.

Statistical analysis
We examined the statistical significance of the association between IPD availability and RCT characteristics and reported RoB dimensions using a $\chi^2$ test or a Fisher exact test for categorical variables, and a Student’s t-test or a Wilcoxon test for continuous variables.

To be comparable to our IPD-available MAs, we calculated global HRs of IPD-unavailable RCTs using the natural logarithm (ln) of the HR and its variance, based on published data, for each MA using a fixed effect model. We calculated the ratio of the global HR of IPD-unavailable RCTs to the global HR of IPD-available RCTs (ie, the ratio of HR (RHR) for each MA. We also calculated a global RHR across MAs and its 95% CI was calculated using a fixed effect model. An RHR less than 1 indicates that IPD-unavailable RCTs present a larger treatment effect estimate than IPD-available RCTs. The heterogeneity of the RHR was assessed with a $\chi^2$ test and the I² statistic. A random effect model was used in cases of statistically significant heterogeneity (p<0.10). Since
RCTs with high RoB are excluded from our published MA (ie, ‘excluded IPD-available RCTs’, noted above) and therefore not included in the global treatment effect of the IPD-available RCTs, a sensitivity analysis of the comparison of treatment effect across IPD-unavailable and IPD-available RCTs was performed without the IPD-unavailable RCTs considered at high RoB for at least one dimension of the RoB tool. Analyses were performed using SAS V.9.3 (SAS Institute, Cary, North Carolina, USA) and R software.37

Patient and public involvement
Patients and the public were not involved in this study.

RESULTS
Three hundred and forty-nine RCTs (73,018 patients included between 1965 and 2010) were eligible for at least one of the 18 MAs. IPD from 60 RCTs (5,890 patients) were unavailable for our published IPD-MAs (figure 1): 32 (53%) because the data were lost by the investigators, 19 (32%) because the investigators could not be contacted, 2 (3%) because the investigators refused to share the RCT data and 7 (12% and 638 patients) for an unknown reason. Of the 289 IPD-available RCTs, 18 RCTs (ie, excluded IPD-available RCTs) were excluded from the published MAs after IPD checking due to major doubts on quality using tools different from the Cochrane RoB tool (figure 1). Reasons for their exclusions are presented in online supplementary etable 2. The main reason was suspicion of biased randomisation (10/18 (56%) excluded RCTs). In the end, 271 RCTs were included in the 18 MAs (figure 1). The 18 excluded IPD-available RCTs have been included only in the analysis of the next paragraph.

Comparison of trial and report characteristics of IPD-available and IPD-unavailable RCTs
Table 1 compares the trial characteristics of the 60 IPD-unavailable RCTs and the 289 IPD-available RCTs. IPD-unavailable RCTs were significantly smaller, more often single centre and non-international. They were also published more often as conference abstracts only and

Figure 1 Flow chart. IPD, individual patient data; MA, systematic review and meta-analysis; OS, overall survival; RCT, randomised controlled trial. *Other MAs are MAs without one IPD-unavailable trial. **Corresponding MAs are MAs including at least one IPD-unavailable trial with extractable HR for OS.
Table 1  Characteristics of randomised trials eligible in one or more of the considered meta-analyses according to the availability of individual patient data (IPD)

| Characteristics of trials | IPD-unavailable (n=60) | IPD-available (n=289) | P values |
|---------------------------|------------------------|-----------------------|----------|
| Number of patients, No (%) |                        |                       |          |
| <50                       | 11 (18)                | 25 (9)                | <0.001*  |
| 50–99                     | 24 (40)                | 63 (22)               |          |
| 100–149                   | 16 (27)                | 45 (16)               |          |
| 150–199                   | 4 (7)                  | 32 (11)               |          |
| 200–249                   | 4 (7)                  | 30 (10)               |          |
| 250–349                   | 1 (2)                  | 43 (15)               |          |
| ≥350 patients             | 0 (0)                  | 51 (18)               |          |
| Date first patient randomised, No (%) |                        |                       |          |
| Before 1980               | 11 (18)                | 32 (11)               | 0.29*    |
| 1980–1984                 | 4 (7)                  | 53 (18)               |          |
| 1985–1989                 | 4 (7)                  | 62 (21)               |          |
| 1990–1994                 | 8 (13)                 | 59 (20)               |          |
| 1995–1999                 | 11 (18)                | 53 (18)               |          |
| 2000–2009                 | 9 (15)                 | 27 (9)                |          |
| Missing                   | 13 (22)                | 3 (1)                 |          |
| Date of publication, No (%) |                        |                       |          |
| Before 1985               | 9 (15)                 | 18 (6)                | 0.24*    |
| 1985–1989                 | 10 (17)                | 37 (13)               |          |
| 1990–1994                 | 6 (10)                 | 52 (18)               |          |
| 1995–1999                 | 8 (13)                 | 51 (18)               |          |
| 2000–2005                 | 17 (28)                | 61 (21)               |          |
| 2005–2014                 | 8 (13)                 | 56 (19)               |          |
| Unpublished               | 2 (3)                  | 14 (5)                |          |
| Number of centres, No (%) |                        |                       |          |
| One or two centres        | 42 (70)                | 116 (40)              | <0.001†  |
| More than two centres     | 14 (23)                | 170 (59)              |          |
| Missing                   | 4 (7)                  | 3 (1)                 |          |
| International trial, No (%) |                        |                       |          |
| Yes                       | 0 (0)                  | 53 (18)               | <0.001†  |
| No                        | 58 (97)                | 236 (82)              |          |
| Missing                   | 2 (3)                  |                      |          |
| Authors’ location, No (%) |                        |                       |          |
| Europe                    | 25 (42)                | 139 (48)              | 0.054‡   |
| North America             | 13 (22)                | 72 (25)               |          |
| Asia                      | 21 (35)                | 51 (18)               |          |
| Central or South America  | 0 (0)                  | 7 (2)                 |          |
| Oceania                   | 0 (0)                  | 4 (1)                 |          |
| Africa                    | 1 (2)                  | 3 (1)                 |          |
| Transcontinental          | 0 (0)                  | 13 (5)                |          |
| Type of publication, No (%) |                        |                       |          |
| Full-text article         | 38 (63)                | 250 (87)              | <0.001†  |
| Conference abstract       | 20 (33)                | 26 (9)                |          |
| Unpublished               | 2 (3)                  | 13 (5)                |          |

Continued
in languages other than English. Although funding of the trial was less often reported in IPD-unavailable trials (p<0.001), it was reported less than 50% of the time for both IPD-available and IPD-unavailable RCTs.

**Comparison of reported quality of IPD-available and IPD-unavailable RCTs**

We found that 27 IPD-unavailable RCTs from 11 MA were published as full-text article in English, and 232 IPD-available RCTs (figure 1). Twelve strata were created based on meta-analysis and time period. They were used to match the 27 IPD-unavailable RCTs with 151 IPD-available RCTs. Because of insufficient RCTs in the IPD-available group in three meta-analyses, the matching was performed in a 1:1 ratio for eight IPD-unavailable RCTs. For 17 others, a random selection of IPD-available RCTs was performed to obtain a 1:2 ratio. For the last two IPD-unavailable RCTs, only four IPD-available RCTs were available. At the end, 46 IPD-available trials were matched. RCTs with a low RoB for random sequence generation and allocation concealment were less frequent, and RCTs with unclear risk were more frequent, in the IPD-unavailable RCTs compared with the IPD-available RCTs (table 2).

Blinding of personnel and participants and blinding of outcome assessment were at low RoB in all cases, because outcome assessed was overall survival, an objective outcome little impacted by lack of blinding. There was no statistically significant difference between IPD-unavailable and IPD-available RCTs in terms of incomplete or missing outcome data.

**Availability of treatment effect estimate for survival in IPD-unavailable RCTs publications**

Only 23/60 IPD-unavailable RCTs (38%), which included 2 434/5 890 patients (41%), had an extractable HR (ie, treatment effect estimate) (see online supplementary etable 3). Extraction was not possible in 37 other trials because of missing or incomplete HR information. RCTs without an extractable HR were more often published as abstracts and less often published in English than RCTs with extractable HR. None of the 20 abstracts had extractable HR (see online supplementary etable 4).

**Comparison of treatment effect across IPD-available and IPD-unavailable RCTs**

The 23 IPD-unavailable RCTs with an available treatment effect measure were included in 10 different MAs. In 9/10 MAs, the RHR was less than 1 (IPD-unavailable RCTs are associated with a greater treatment effect estimate than IPD-available RCTs). Across all MAs, IPD-unavailable RCTs observed a 14% greater treatment effect than what was observed in IPD-MAs (global RHR=0.86; 95% CI 0.75 to 0.98, p=0.025). There was no significant heterogeneity (p=0.24) among meta-analysis results (figure 2). After exclusion of the 11 IPD-unavailable RCTs considered at high RoB (the sensitivity analysis), the global RHR was 1.01 (0.84–1.21).

**DISCUSSION**

In the 18 IPD-MAs coordinated by Gustave Roussy, IPD from 60/349 (17%) eligible trials and 5890/66 928 (8%) eligible patients were not made available to us. Combining IPD with aggregate data could lead to biased estimates of a treatment effect if the trials with aggregate data only (IPD-unavailable) are different from IPD-available trials. Regrettably, based on reported information, IPD-unavailable trials are different in both their characteristics and RoB. IPD-unavailable RCTs were smaller, more often single centre and non-international. They were also published more often as conference abstracts.
only and in languages other than English. We found that the RoB in IPD-unavailable RCTs was greater than the RoB in IPD-available RCTs.

In our study, 37/60 (62%) of IPD-unavailable RCTs did not contribute to the summary effect estimate because we could not obtain an HR. Thus, we do not know whether the results from these trials are higher, lower or the same as the estimate we obtained for the 23 trials with HR data. We are surprised that overall survival could not be assessed for so many trials, given that it is one of the most important cancer trial endpoints.

For the other 23/60 RCTs, we inferred from the RHR that IPD-unavailable RCTs are associated with a larger treatment effect estimate than IPD-available RCTs, that is, a relative increase of 14% of treatment effect on overall survival was observed in the IPD-unavailable trials, even though our MAs were in the field of oncology in which treatment effect are often small. When, in a sensitivity analysis, we removed 11/23 trials assessed at high RoB from the analysis, RHRs for IPD-unavailable RCTs compared with IPD-available RCTs implied no discernible difference between the two effect estimates.

In another study of 31 IPD-MAs, not including any of ours, the median proportion of IPD-available RCTs across MAs was 91% (range 60%–100%) compared with 85% (range 71%–100%) in our study. In our study, IPD-availability is lower in RCTs published as conference abstracts only than in those published as full-text article (table 1). Then, IPD-availability rate may depend on whether the ‘grey literature’ (ie, meeting abstracts and communications) has been searched or not. According to Ahmed et al, only 9/31 meta-analyses mentioned grey literature compared with 100% in our MAs. This may explain our lower rate of available data.

Should we combine IPD-unavailable information (aggregated HRs from individual trials) with IPD-available information? Although some data are better than no data, including aggregated data in the MAs would not have solved entirely the problem of missing data. Only 23/60 (38%) of the IPD-unavailable trials provided a meta-analyzable treatment effect estimate. Other treatment effect estimates other than HR, such as comparison of survival rates or comparison of median survival times, were not considered since they are less appropriate for survival endpoints. In addition, IPD-unavailable trials had different characteristics and a higher RoB. This is not to say that IPD-available trials are never at a RoB. Many trials in our sample, including IPD-available trials, did not report sufficient information to evaluate the RoB properly, as shown previously by others. What we do not know is whether lack of clarity in the report reflects poorer methodological quality of the trial. By checking IPD and exchanging with investigators, it was possible to confirm or complete the published information for the IPD available trials. Not combining IPD-unavailable information with IPD-available information is open to selection bias. In our experience, when an IPD-MA and an aggregated MA were performed at the same time, there were more trials and patients included in the IPD-MA, and then the selection bias may be lower for IPD-MA than for the corresponding aggregated MA.

Our study has several strengths. First, it is the first study, to our knowledge, to study the characteristics of IPD-unavailable RCTs compared with IPD-available RCTs and their treatment effects. Second, the standardised methods used both in selecting the RCTs to be included in the MA and in the methods of our MA assure minimisation of bias and meta-bias. Third, we examined 18 MAs and 349 eligible RCTs, a large sample that contributes robustness to our results. Last, our results are consistent from one MA to the other.

The main limitation of this study is that we included only MAs with overall survival data, performed by our group, in oncology. But based on our collaboration with other groups performing IPD-MA, and our participation in the development of the guidelines for IPD-MAs, we think our IPD-MA methods correspond to those used by other teams. Then, our results would be applicable to

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Table 2  Quality of randomised trials by IPD availability, as obtained from English language full-length publications

| Dimensions of the Cochrane risk of bias tool | RCTs in MA | IPD- unavailable (n=27) | IPD- available* (n=46) | P values |
|---------------------------------------------|-----------|-------------------------|------------------------|----------|
| Random sequence generation, No (%)         |           |                         |                        |          |
| Low risk                                   | 9 (33)    | 32 (70)                 | 0.007†                 |          |
| High risk                                  | 1 (4)‡    | 0 (0)                   |                        |          |
| Unclear                                    | 17 (63)   | 14 (30)                 |                        |          |
| Allocation concealment, No (%)             |           |                         |                        |          |
| Low risk‡‡                                 | 2 (7)     | 17 (37)                 | 0.006§                 |          |
| High risk                                  | 0 (0)     | 0 (0)                   |                        |          |
| Unclear                                    | 25 (93)   | 29 (63)                 |                        |          |
| Blinding of personnel and participants, No (%) |               |                         |                        |          |
| Low risk¶                                 | 27 (100)  | 46 (100)                | –                      |          |
| Blinding of outcome assessment, No (%)     |           |                         |                        |          |
| Low risk¶                                 | 27 (100)  | 46 (100)                | –                      |          |
| Incomplete outcome data (attrition bias), No (%) |               |                         |                        |          |
| Low risk††                                | 7 (26)    | 18 (39)                 | 0.21§                  |          |
| High risk                                  | 11 (41)   | 10 (22)                 |                        |          |
| Unclear                                    | 9 (33)    | 18 (39)                 |                        |          |

*Sample paired with unavailable IPD trials on the first inclusion period and the meta-analysis which included them, with a 2:1 ratio.
†Chi square test.
‡Classified at high risk because of a randomization by pairs.
§Fisher exact test.
¶Outcome examined was overall survival.
IPD, individual patient data; RCT, randomised clinical trial.
other topics where overall survival is the main outcome, such as cardiovascular disease. Other studies are needed to confirm applicability, however. Another limitation is the lack of evaluation of the impact of the exclusion of 18 IPD available RCTs on treatment effect due to the loss of the data of 7 of them. But results similar to the IPD-unavailable RCTs, in particular those with high RoB, are expected since the main reason for exclusion was the high RoB for randomisation. Lastly, the small number of recent trials in our sample could affect reporting of trials characteristic, RoB and treatment effects. However, unavailable IPD are still observed in current MAs our group is performing, and the proportion of unavailable IPD has not varied with the date of randomisation of the first patient in each RCT.

Our results indicate that IPD-unavailable RCTs compared with IPD-available RCTs report a higher treatment effect.

There are multiple reasons that may explain the difference in the observed treatment effect between aggregated IPD-unavailable data and IPD-available data RCTs.► IPD-unavailable RCTs were significantly smaller than IPD-available RCTs and if they were included in an MA, they could contribute to a biased treatment effect estimate. Indeed, it has been previously observed on aggregated data MAs that small studies tended to show a larger effect size than bigger studies19,50, the reason for this overestimation remains unclear.
► As we saw in our sensitivity analysis, IPD-unavailable RCTs could be at higher RoB than IPD-available RCTs, and omission of trials at high RoB means eliminating the overestimation of treatment effect.
► We saw that characteristics of IPD-unavailable RCTs compared with excluded IPD-available RCTs (online supplementary etable 2) share common traits: they tend to be non-international single-centre RCTs, mainly Asian, and are frequently not published in English.
► It has been previously observed that trials with higher RoB or lower methodological quality tend to report greater effect sizes.51 52
► Since IPD-available data sets are checked, completed and/or corrected if needed by the coordinating group in collaboration with the investigators, and IPD-unavailable aggregated data are used as published, the two sources of data are different. For example, follow-up was updated for IPD-available but not IPD-unavailable RCTs, affecting our MAs, and the shorter follow-up in IPD-unavailable trials may also explain in part the higher treatment effect observed.

Figure 2  Forest plot for ratio of HR (HR in individual patient data (IPD)-unavailable trials reported to HR in IPD meta-analyses (MAs)). Note that for 5 out of 18 MAs (MA-2, MA-11, MA-12, MA-13 and MA-15), data from all trials were available; for three other, it was not possible to extract an HR from the publications of the unavailable trials (MA-4, MA-5 and MA-9). Analysis on the MA with available IPD did not include trials excluded for quality reasons. For the available IPD MA, results are based on the corresponding publication and 13 randomised controlled trials (RCTs) with IPD were excluded for quality reasons. The MAs and the status of the corresponding RCTs are described in online supplementary etable 1. CI, confidence interval; HR, hazard ratio; pts, patients; RHR, ratio of HR.
We cannot exclude the possibility that at least some of IPD-unavailable information is fabricated or fraudulent since the data set is not shared.

When aggregated IPD-unavailable information is excluded from the MA, the treatment effect is conservatively estimated and the treatment effect is closer to no effect.

It is important to include all eligible RCTs in any MA. It is particularly important in cases such as ours, when the IPD-available RCTs appear to be different from the IPD-unavailable RCTs, in their characteristics, RoB, and HRs for survival. Because 60/349 (17%) of all RCTs eligible for our 18 MAs were IPD-unavailable trials, data sharing appears of utmost importance to improve reliability of MA results. Several initiatives from profit and non-profit organisation to share trial IPD or to promote it are ongoing; some of them already allowing facilitated access to trial databases. In spite of the urgent need for efficient sharing data system providing access to trial data from around the world, it will take some time before user-friendly systems that respect the rights of all stakeholders are in place. Access to the IPD from older trials may never be available.

Even if IPD becomes widely available, we recommend at least the following, relevant to our findings:

- Mandatory systematic review (involving a comprehensive search for eligible trials) before meta-analysis;
- A comprehensive search for published trial results should include multiple bibliographic databases and reports in all languages;
- Central registration of all trials is essential for an appropriately comprehensive search of all trials (published and unpublished) eligible for a systematic review;
- Systematic reviews should include evaluation of individual trial RoB, since biased trials can lead to a biased meta-analysis and use tools adapted to the type of data (ie, aggregated data or IPD). Assessment of a trial RoB can be facilitated by making trial protocols public;
- Systematic reviewers should consider the potential impact of unpublished trials and selective reporting of trial information, and obtain the needed unpublished information if possible;
- Meta-analysts should describe the eligible trials they found that have unavailable data (number of trials, number of patients, reasons for unavailability, precise reference). For IPD-MA, the amount of unavailable IPD and among them, proportion of summary data available and their impact on the results should be discussed. If the authors did not submit the data suggested, reviewers and editors should request them.
- Consolidated Standards of Reporting Trials reporting guidelines for trials should be endorsed by all journals and adhered to by all authors to ensure that meta-analyzable data are provided in all full and abstract publications.

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