Introduction

Community-acquired pneumonia (CAP) is a major cause of morbidity, mortality and health care burden globally.\textsuperscript{1,2} Certain population groups are more susceptible, particularly the elderly and those with chronic conditions such as COPD.\textsuperscript{3,4} The prevalence of COPD increases with age and is estimated to affect ~210 million people globally, and therefore this population contributes significantly to the overall disease burden of CAP.\textsuperscript{5,6}
The clinical course of COPD is interspersed by episodes of symptomatic deterioration, termed exacerbations. CAP and exacerbations can be difficult to distinguish clinically, and there remains no clear consensus on whether they represent distinct events or are in fact a continuum. Exacerbations are highly seasonal, with increasing rates seen in the winter months, adding further to the pressures already experienced by overstretched health care providers at this time of year. The winter peak coincides with the increased activity of viral infection, and a causal relationship has been proposed. Although exacerbations are more frequent in winter, it is not known whether COPD-associated CAP has a similar seasonal pattern. In Europe, hospitalizations due to CAP are increasing, with COPD a frequently encountered comorbidity. Reports suggest that patients with COPD hospitalized with CAP present with more severe symptoms and have longer hospital stays and increased need for ICU admission. CAP is common in the natural history of COPD, with a much higher incidence compared to that of the adult general population. Furthermore, in a UK-based audit, 18% of COPD patients requiring hospitalization for acute respiratory illness had consolidation evident on chest X-ray, diagnostic of CAP. Recent reports have highlighted an increased risk of CAP associated with inhaled corticosteroid (ICS) use in COPD patients, underlining the need for stratification of patients for ICS use to include an estimation of this risk. Therefore, COPD-associated CAP is particularly relevant considering the aging global population and will have important implications for future antibiotic usage and health care resources.

Despite the greater incidence of CAP in COPD compared to the adult general population, it remains a relatively infrequent event compared to exacerbations. Consequently, the true burden of CAP in COPD has been largely overlooked and limited to predominantly cross-sectional analyses of hospitalized patients. Analysis of large population databases can provide a “real-world” perspective of incidence patterns, health care usage and risk factors of CAP in COPD and, thus, an understanding of the likely triggers, susceptibilities and socioeconomic burden of what still remains a poorly understood event in the natural history of many COPD patients. This is key to guiding future, targeted research, focusing on preventative strategies in vulnerable patient groups.

Linked primary and secondary health care records have the potential to improve the accuracy of disease incidence estimates compared to stand-alone records. Therefore, using the Hampshire Health Record Analytical (HHRA) database, an anonymized database of linked health care records covering ~75% of primary care practices in Hampshire, Southern England, we sought to describe the seasonal patterns, hospital admissions and risk factors for CAP, among patients with COPD.

Methods
Study design and population
A retrospective, observational cohort design was used. Patients were defined as having COPD and eligible for inclusion if they had ≥1 COPD-specific Read code (standard UK classification of clinical terms for electronic coding) in their primary care record at any time prior to January 1, 2010, and were aged ≥40 years at the start of the study. Follow-up ran for 5 years from January 1, 2010, until December 31, 2014.

The HHRA database was used for data collection (Supplementary material). Briefly, it represents a large electronic database of individual patient-anonymized primary and secondary care records, providing a single source of linked routine health data for patients living in Hampshire, Southern England. At the time of the study, the HHRA contained records covering ~1.4 million patients from >140 primary care practices.

Use of the HHRA database for research is regulated by the Hampshire Health Record Information Governance Group (HHRIGG) and National Health Service (NHS) South, Central and West Commissioning Support Unit. NHS data systems entitle patients to opt-out of consenting to storage and use of personal information for analytical purposes. Furthermore, as the HHRA data are anonymized, it is not possible to identify patients; as such, and as confirmed by local research governance (University of Southampton Research, Integrity and Governance Team), formal ethical approval and written informed consent from all participants were not required for this study.

Demographics and clinical factors
Data were collected on sex, age, weight, height, body mass index (BMI) and disease severity by Global Initiative for Obstructive Lung Disease (GOLD) stage at the most recently recorded time point prior to the study start date, within a pre-assigned limit of 2 years. Comorbidities were identified using Read code lists and included as present if recorded at any time prior to the study start. Usage of concurrent therapies for COPD was included if ≥1 prescription was evident in the 6 months prior to the study start, with the exception of oral corticosteroids (OCSs), where frequent use was defined as ≥4 prescriptions in the year prior to the study start. Code lists for all the covariates included are available on request.
COPD-associated CAP
CAP episodes were identified either from a relevant International Classification of Diseases Version 10 (ICD-10) code in the primary diagnostic position following an inpatient hospital stay or emergency department (ED) attendance or from a physician Read-coded diagnosis of CAP. In order to avoid capturing multiple-coded events for one illness as different episodes, single-coded CAP events were grouped into clusters. Clusters were defined as the number of CAP events occurring within a continuous 70-day period, with the date of the first event regarded as the index date. A period of at least 2 event-free weeks following CAP episodes was required in order to avoid labeling prolonged recording of events for the same illness, as new episodes.

COPD exacerbations
Exacerbations were defined as either, 1) a physician Read-coded diagnosis of exacerbation, 2) the prescription of specific antibiotics or OCS on the same day or up to 7 days before or after a COPD diagnostic or symptom Read code, or 3) a relevant respiratory ICD-10 code in the primary or secondary diagnostic positions, following an inpatient hospital stay or ED attendance. A continuous 21-day period defined an episode, with 2 event-free weeks required to separate exacerbations.

Seasonality
The month of the index event (CAP or exacerbation) date was taken as the month of episode occurrence. Winter was defined as December to February and summer as June to August.

Statistical analysis
Analyses were performed using SPSS version 22. CAP and exacerbation rates were calculated per month, adjusting for mortality, and displayed over time. Categorical variables are presented as number (percentage). Chi-square tests were used to compare data between groups with and without a diagnosis of CAP. Multivariate logistic regression was then used to identify adjusted odds ratios (ORs) and 95% confidence intervals (95% CIs) for having ≥1 CAP episode versus none, in patients with complete data, who either survived until the end of the observation period or died following CAP. Covariates reaching significance from the bivariate analysis were included in the multivariate model. To identify risk factors for hospital admission during CAP, we used STATA to conduct conditional logistic regression, treating subject as a random effect. This accounted for the over-representation of some patients (ie, those experiencing >1 CAP). A P-value of <0.05 was considered statistically significant.

Results
Population characteristics
The study cohort consisted of 14,513 patients with a diagnosis of COPD. The mean ± standard deviation (SD) age at the start of the study period was 70.3±10.8 years; 53.6% were male and 33.5% were current smokers (Table 1). Of the cohort in which lung function data were available (n=10,358), mean ± SD forced expiratory volume in 1 second (FEV1) %predicted was 59.5±19.8. Overall, exacerbations were more prevalent than CAP. The incidence rate for CAP was 37.6 per 1,000 person-years, with 13.4% (n=1,938) experiencing at least one CAP, and of these, 18.8% (n=364) had more than one.

Seasonality of CAP, exacerbations and hospital admissions
Figure 1A and B displays the total monthly rates of CAP and exacerbations across 5 years, while Figure 1C and D displays the rates averaged into 1 calendar year. A clear seasonal pattern is evident with highest rates of CAP in the winter, especially December, compared to other periods of the year. Median monthly CAP rates were 4.0 per 1,000 person-months in winter versus 2.8 in summer (P=0.003). Similarly, for exacerbations, higher rates occurred in winter, with lower rates in summer (83.6 versus 60.5 per 1,000 person-months, respectively, P<0.001).

A total of 2,413 CAP episodes were identified over 5 years. Trends in hospital admissions associated with CAP closely followed the seasonal distribution of total CAP episodes, with 75.6% of CAP associated with hospital admission, compared to only 22.1% of exacerbations.

Risk factors for CAP
Using multivariate logistic regression modeling and controlling for potential confounders, factors independently predicting CAP risk were identified (Table 2). The odds of experiencing CAP increased with age, with patients aged 60–79 and ≥80 years having higher odds than those aged 40–59 years (OR 1.67, 95% CI 1.30–2.16 and OR 4.10, 95% CI 3.05–5.49, respectively). BMI was also strongly associated with the odds of CAP, with those underweight having higher odds (OR 1.82, 95% CI 1.27–2.62) than those of normal weight. Conversely, the odds of CAP in overweight and obese patients was 30% lower (OR 0.73, 95% CI 0.61–0.87 and OR 0.70, 95% CI 0.58–0.85, respectively) than that in patients of normal weight. Increasing grade of airflow limitation was associated with higher odds of CAP, with GOLD stage IV associated with almost three
times the odds compared to GOLD stage I (OR 2.86, 95% CI 2.04–4.09).

Additional factors independently associated with increased odds of CAP were a prior history of CAP and comorbid diseases, including ischemic heart disease (IHD), diabetes and chronic kidney disease (CKD).

### Exacerbation frequency and CAP risk

A history of frequent exacerbations, defined as ≥2 exacerbations in the year prior to the study start date, appears to be predictive of future CAP. Frequent exacerbators had almost 30% higher odds of having CAP over the study period (OR 1.29, 95% CI 1.08–1.54), compared to infrequent exacerbators.

### Table 1 Patient demographics and baseline clinical characteristics of the total COPD cohort, by presence or absence of a CAP diagnosis over the 5-year study period

| Characteristic                        | Total  | No CAP          | ≥1 CAP            | P-value |
|---------------------------------------|--------|-----------------|-------------------|---------|
| Number of patients                    | 14,513 | 12,575 (86.6)   | 1,938 (13.4)      | 0.145   |
| Sex (male)                            | 7,775  | 6,707 (53.3)    | 1,068 (53.1)      | <0.001  |
| Age (years)                           |        |                 |                   |         |
| 40–59                                 | 2,347  | 2,171 (73.5)    | 176 (6.1)         |         |
| 60–79                                 | 9,044  | 7,839 (63.1)    | 1,105 (57.0)      |         |
| ≥80                                   | 3,122  | 2,465 (78.9)    | 657 (33.0)        | 0.027   |
| Smoking status (n=12,972)             |        |                 |                   |         |
| Current                               | 4,859  | 4,255 (37.8)    | 604 (35.1)        |         |
| Ex/never                              | 8,113  | 6,994 (62.2)    | 1,119 (64.9)      |         |
| BMI, kg/m² (n=9,235)                  |        |                 |                   | <0.001  |
| Underweight (<18.5)                   | 340    | 275 (3.4)       | 65 (5.5)          |         |
| Normal weight (18.5–24.9)             | 2,911  | 2,483 (80.9)    | 428 (36.1)        |         |
| Overweight (25.0–29.9)                | 3,247  | 2,874 (88.4)    | 373 (31.4)        |         |
| Obese (≥30)                           | 2,737  | 2,416 (88.4)    | 321 (37.0)        |         |
| Prior CAP                             | 358    | 253 (2.0)       | 105 (5.4)         |         |
| Exacerbation frequency                |        |                 |                   | <0.001  |
| Frequent (≥2 per year)                | 2,496  | 2,004 (15.9)    | 492 (25.4)        |         |
| Infrequent (<2 per year)              | 12,017 | 10,571 (84.1)   | 1,446 (74.6)      |         |
| Pharmacotherapy                       |        |                 |                   |         |
| ICS use                               | 7,806  | 6,348 (50.5)    | 1,112 (57.9)      | <0.001  |
| Frequent OCS use                      | 1,202  | 925 (7.4)       | 277 (14.3)        | <0.001  |
| LABA use                              | 6,376  | 5,353 (42.6)    | 1,023 (52.8)      | <0.001  |
| LAMA use                              | 4,046  | 3,323 (26.4)    | 723 (37.3)        | <0.001  |
| Prior pneumococcal vaccine            | 4,790  | 4,167 (33.1)    | 623 (32.1)        | 0.388   |
| Comorbidity                           |        |                 |                   |         |
| IHD                                   | 2,853  | 2,347 (18.7)    | 506 (26.1)        | <0.001  |
| Heart failure                         | 931    | 739 (6.9)       | 192 (9.9)         | <0.001  |
| Hypertension                          | 6,043  | 5,146 (40.9)    | 897 (46.3)        | <0.001  |
| Diabetes mellitus                     | 1,978  | 1,622 (12.9)    | 356 (18.4)        | <0.001  |
| Peripheral vascular disease           | 852    | 692 (5.5)       | 160 (8.3)         | <0.001  |
| Hyperlipidemia                        | 2,514  | 2,203 (17.5)    | 338 (17.4)        | 0.933   |
| Cerebrovascular disease               | 1,397  | 1,115 (8.9)     | 282 (14.6)        | <0.001  |
| Dementia                              | 306    | 245 (1.9)       | 61 (3.1)          | 0.001   |
| Asthma                                | 6,989  | 5,982 (47.6)    | 1,007 (52.0)      | <0.001  |
| GORD                                  | 1,423  | 1,198 (8.9)     | 225 (11.6)        | 0.004   |
| Anxiety or depression                 | 5,380  | 4,643 (36.9)    | 737 (38.0)        | 0.348   |
| Osteoporosis                          | 955    | 780 (6.2)       | 174 (9.0)         | <0.001  |
| CKD                                   | 2,263  | 1,836 (14.6)    | 427 (22.0)        | <0.001  |
| Connective tissue disease             | 350    | 280 (2.2)       | 70 (3.6)          | <0.001  |
| GOLD staging (n=10,358)               |        |                 |                   | <0.001  |
| I                                      | 1,524  | 1,387 (15.4)    | 137 (10.0)        |         |
| II                                     | 5,396  | 4,795 (53.3)    | 601 (43.9)        |         |
| III                                    | 2,803  | 2,299 (25.6)    | 504 (36.8)        |         |
| IV                                     | 635    | 509 (5.7)       | 126 (9.2)         |         |

**Notes:** Data are presented as frequencies (%). *P*-values from chi-square test.

**Abbreviations:** BMI, body mass index; CAP, community-acquired pneumonia; CKD, chronic kidney disease; GOLD, Global Initiative for Obstructive Lung Disease; GORD, gastroesophageal reflux disease; ICS, inhaled corticosteroid; IHD, ischemic heart disease; LABA, long-acting beta-agonist; LAMA, long-acting muscarinic antagonist; OCS, oral corticosteroid.
exacerbators. In addition, a greater proportion of frequent exacerbators experienced recurrent CAP over 5 years, compared with infrequent exacerbators (5.1% versus 2.0%, \( P<0.001 \)) (Figure 2).

**Risk factors for hospital admission associated with CAP**

Conditional logistic regression, accounting for subject effect, was used to examine associations between patient characteristics and the odds of hospital admission during CAP episodes (n patients =1,368, n CAP =1,721). Increasing age and grade of airflow limitation were independently associated with increased odds of hospital admission during CAP (Table 3).

**All-cause mortality**

A total of 2,877 deaths occurred over the 5-year study period. All-cause mortality followed a similar seasonal distribution to CAP and exacerbations (Figure 3); 880 (30.6%) deaths occurred in winter compared to 602 (20.9%) in summer.
Table 2 Adjusted ORs and 95% CIs for predictors of 5-year CAP risk

| n=6,805 | OR    | 95% CI      | P-value |
|---------|-------|-------------|---------|
| Sex     |       |             |         |
| Female  | Reference |           |         |
| Male    | 1.182 | 1.015–1.376 | 0.031  |
| Age (years) |       |             |         |
| 40–59   | Reference |           |         |
| 60–79   | 1.674 | 1.296–2.161 | <0.001 |
| ≥80     | 4.096 | 3.053–5.494 | <0.001 |
| Smoking status |       |             |         |
| Ex/never | Reference |           |         |
| Current | 1.316 | 1.121–1.544 | 0.001  |
| BMI, kg/m² |       |             |         |
| Underweight (<18.5) | 1.823 | 1.267–2.624 | 0.001  |
| Normal weight (18.5–24.9) | Reference |           |         |
| Overweight (25–29.9) | 0.729 | 0.610–0.872 | 0.001  |
| Obese (≥30) | 0.701 | 0.576–0.853 | <0.001 |
| Prior CAP |       |             |         |
| Exacerbation frequency |       |             |         |
| Infrequent (≤2 per year) | Reference |           |         |
| Frequent (≥2 per year) | 1.291 | 1.080–1.543 | 0.005  |
| Pharmacotherapy |       |             |         |
| ICS use | 1.229 | 1.001–1.508 | 0.049  |
| Frequent OCS use | 1.934 | 1.539–2.431 | <0.001 |
| LABA use | 1.136 | 0.936–1.379 | 0.196  |
| LAMA use | 1.398 | 1.199–1.628 | <0.001 |
| Comorbidities |       |             |         |
| IHD     | 1.312 | 1.102–1.562 | 0.002  |
| Heart failure | 1.415 | 1.065–1.881 | 0.017  |
| Hypertension | 1.081 | 0.929–1.258 | 0.315  |
| Diabetes mellitus | 1.736 | 1.438–2.097 | <0.001 |
| Peripheral vascular disease | 1.364 | 1.044–1.784 | 0.023  |
| Cerebrovascular disease | 1.325 | 1.050–1.672 | 0.018  |
| Dementia | 2.219 | 1.271–3.877 | 0.005  |
| Asthma  | 1.049 | 0.901–1.220 | 0.538  |
| GORD    | 1.179 | 0.948–1.467 | 0.138  |
| CKD     | 1.322 | 1.093–1.600 | 0.004  |
| Osteoporosis | 1.173 | 0.882–1.559 | 0.273  |
| Connective tissue disease | 1.688 | 1.110–2.565 | 0.014  |
| GOLD stage |       |             |         |
| I       | Reference |           |         |
| II      | 1.296 | 1.018–1.649 | 0.035  |
| III     | 2.239 | 1.735–2.891 | <0.001 |
| IV      | 2.861 | 2.003–4.085 | <0.001 |

Notes: Multivariate logistic regression model of 6,805 patients with complete data, who either survived until the end of the 5-year observation period or died during the observation period following a CAP episode and were not lost to follow-up. For comorbidities and pharmacotherapy, not having the comorbidity and not using the drug were taken as the reference.

Abbreviations: BMI, body mass index; CAP, community-acquired pneumonia; CI, confidence interval; CKD, chronic kidney disease; GORD, gastroesophageal reflux disease; ICS, inhaled corticosteroid; IHD, ischemic heart disease; LABA, long-acting beta-agonist; LAMA, long-acting muscarinic antagonist; GOLD, Global Initiative for Obstructive Lung Disease; OCS, oral corticosteroid; OR, odds ratio.

Discussion

This is the first population-based study showing the seasonal trends of CAP and exacerbations among people with COPD, using linked primary and secondary care data sources. The incidence rate of 37.6 CAP cases per 1,000 person-years compares to that of 22.4 reported by Mullerova et al4 in COPD patients and 7.99 reported by Millett et al3 of adults aged ≥65 years, not limited to COPD.

A total of 13.4% of our study population had CAP over 5 years, with almost a fifth of these having more than one episode. CAP occurrence is to some extent a predictable phenomenon and tracks with exacerbation frequency. Moreover, there are potentially modifiable factors such as BMI and concurrent medication use, which could form the basis for personal use only.

Table 3 Adjusted OR and 95% CI for hospital admission with CAP

| n=1,721 | OR    | 95% CI      | P-value |
|---------|-------|-------------|---------|
| Sex     |       |             |         |
| Female  | Reference |           |         |
| Male    | 1.104 | 0.779–1.565 | 0.576  |
| Age (years) |       |             |         |
| 40–59   | Reference |           |         |
| 60–79   | 1.875 | 1.050–3.348 | 0.034  |
| ≥80     | 4.559 | 2.341–8.878 | <0.001 |
| ICS use | 1.428 | 0.973–2.097 | 0.069  |
| Prior CAP |       |             |         |
| Exacerbation frequency |       |             |         |
| Infrequent (≤2 per year) | Reference |           |         |
| Frequent (≥2 per year) | 1.143 | 0.780–1.676 | 0.494  |
| GOLD stage |       |             |         |
| I       | Reference |           |         |
| II      | 2.325 | 1.307–4.134 | 0.004  |
| III     | 2.784 | 1.530–5.068 | 0.001  |
| IV      | 4.536 | 1.960–10.501 | <0.001 |

Note: Multivariate conditional logistic regression model, including the patient as a random effect, of 1,721 CAP episodes (occurring in 1,368 patients) with complete data.

Abbreviations: CAP, community-acquired pneumonia; CI, confidence interval; ICS, inhaled corticosteroid; GOLD, Global Initiative for Obstructive Lung Disease; OR, odds ratio.
of targeted risk-reducing strategies. Greater understanding of the underlying mechanisms is therefore required to establish why some patients are more susceptible to CAP than others.

CAP occurred more commonly in the winter, compared to other times of the year. Similarly for exacerbations, and in line with findings from the Towards a Revolution in COPD Health (TORCH) study, a clear and consistent seasonal pattern from year to year was seen. Importantly, our study population differs from the TORCH study, as it benefits from providing a true insight into real-world population trends across all severity of COPD, as compared to a highly selected patient cohort with more severe disease.

The majority of reports to date suggest that CAP in COPD is caused primarily by bacterial infection. However, the microbiological profile of CAP in adult patients without COPD has a proven seasonal basis with an important role attributed to viral infection. Holter et al demonstrated that CAP occurring in the winter and spring is associated with a greater frequency of bacterial–viral co-infection, supporting the results from an earlier study by Lieberman et al, showing greater frequencies of respiratory viral infection ascribed to CAP in these seasons. There is also a large body of evidence supporting the role of influenza infection and secondary bacterial pneumonia; yet, this finding is not apparent in reports for CAP in COPD. A separate analysis of our data (not shown) for each individual influenza season during the 5-year period indicated no strong associations between reduced CAP rates and receipt of influenza vaccination. Certainly, our data show a greater rate of CAP in the winter compared to other times of the year, suggesting that respiratory viral infection may have an important causative role. Indeed, environmental factors, notably lower temperatures and changes in relative humidity during winter, may favor the survival of respiratory viruses and increase host susceptibility to infection. It was not until relatively recently that viral infection and cold symptoms, a marker of putative viral infection, were established as important determinants of COPD exacerbations, leading to worsening symptoms and prolonged recovery. With the proven seasonality of exacerbations and now CAP in COPD, viral infection must be considered as a key early determinant in the sequence of events leading to both exacerbations and CAP in COPD patients.

Like other reports, we found that increasing age, poorer lung function, lower BMI and comorbid disease, particularly IHD, diabetes and CKD, were all independently associated with future CAP risk. It is already known that the elderly, especially those aged >65 years, are particularly susceptible to developing CAP. The mechanisms are complex, but age-influenced immune system dysregulation, predisposing to infection risk is likely to be of importance. On a population-based level, physicians may be more likely to diagnose respiratory infection in older, frailer patients, and yet, we found no observed influence of age on the risk of exacerbation (data not shown). Like previous studies, we show the association between lower BMI and increased CAP risk and higher BMI with reduced risk, compared to normal weight patients. This perceived protective effect of a higher BMI has also been reported previously in studies not limited to COPD patients. Yet, despite the underlying reasons remaining unclear, it could still provide a potentially useful target for risk-reducing interventions.

A susceptibility to frequent exacerbations is now a well-established COPD phenotype and has been shown to be fairly consistent over time. Frequent exacerbators are more susceptible to viral infection and have higher levels of airway inflammation at stable state, both proposed as possible underlying mechanisms. We have established prior exacerbation frequency as an independent risk factor for CAP, with an increased proportion of frequent exacerbators experiencing recurrent CAP. However, a prior history of frequent exacerbations does not appear to influence the likelihood of more severe (ie, hospitalized) CAP events. Furthermore, and in line with a previous report, we show that a prior history of CAP is a risk factor for future CAP. Whether the mechanisms underlying exacerbation frequency and CAP susceptibility are the same requires further research.

As anticipated, our study re-confirms ICS use as an independent risk factor for future CAP, with a further trend evident between ICS use and increased odds of hospital admission associated with CAP. Interestingly, long-acting muscarinic
antagonist (LAMA) use was independently associated with increased odds of CAP. Published data have shown that in COPD patients with severe disease (FEV₁ < 50% predicted), LAMA use appears to confer less risk of CAP than ICS use.17 However, data pertaining to CAP risk with LAMA use in large population analyses of COPD patients and importantly across all disease severities are currently lacking.

The burden of respiratory-related health care utilization and hospital attendances, especially for COPD and CAP, is substantial, with the winter period being particularly pressured.33,34 In line with an increased rate of CAP in winter, we have shown that the number of hospital admissions with CAP also increases substantially, proportionally much more so than with exacerbations. Patients who required hospital admission were older and had greater severity of underlying COPD, representing a vulnerable group and those most likely to impose a greater resource burden.35 More than 75% of CAP episodes were associated with hospital admission, which is much higher than most reports for the adult general population6,7 and probably indicates a lower threshold for admitting COPD patients to hospital, or a manifestation of more severe symptoms. Regardless, with the aging global population and increasing prevalence of COPD, CAP will become a growing future problem in this chronic disease group.

We acknowledge that this study has limitations. We were unable to validate CAP diagnoses with independent assessment of chest X-ray findings, as this was a large population database containing anonymized data in which radiological features were assessed by the attending physician. However, Skull et al36 have reported that ICD-10 codes can be a valid method of retrospectively identifying CAP from medical records. A large proportion of CAP in our study was associated with hospital admission and based on hospital discharge diagnoses, and therefore it is highly likely that a chest X-ray was performed in these cases, confirming the diagnosis.

With regard to our study population, we included patients with a Read-coded diagnosis of COPD and did not require spirometric confirmation for study inclusion. However, this method has been shown to accurately identify COPD patients from UK primary care records with spirometry only marginally improving the accuracy.39

Conclusion

Like exacerbations, COPD-associated CAP is highly seasonal, with an increased burden observed over winter. Despite occurring less frequently in the natural history of COPD than exacerbations, its relative impact on health care utilization is considerable, with the most vulnerable patients often at greatest risk of both events. In addition, establishing common risk factors may allow for a prospective estimation of a patient’s risk of future CAP, aiding clinicians in treatment and management decisions. Studies, focusing on preventative strategies such as interventions to reduce risk and improve outcomes, including nutritional interventions, treatment stratification and new and improved vaccination programs, are key to tackling the increasing future burden of CAP in at-risk patients, such as those with COPD.

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Disclosure

The authors report no conflicts of interest in this work.

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**Supplementary material**

**Methods – data source**

**Hampshire Health Record (HHR) and electronic health care data sets**

The HHR is an electronic shared care record for people living in Hampshire that provides a single source of linked routine health data for individual patients, collected in “real time”, and links data from general practitioner practices and hospitals and recently includes some data from social care. It is able to combine data from many different IT systems. It is not part of the National Summary Care Record but rather is an arrangement local to Hampshire, storing data from a wide range of practices that have “opted in” to the data-sharing scheme. The HHR is used by clinicians to share information and aid clinical decision-making and is currently supported with data supplied by many organizations in the South of England, including local primary care practices and providers of secondary, community and social care.

**Hampshire Health Record Analytical (HHRA) database**

The HHRA database is a separate electronic database created for research and analysis to support health improvement and planning. At the time of this study, it included data from >140 practices across Hampshire, covering ~1.4 million patients. It shares some of the health data contained within the HHR but in anonymized form to protect patients’ identity. The HHR and the HHRA are hosted and maintained by NHS South, Central and West Commissioning Support Unit (CSU). The Governance body is the HHR Information Governance Group (HHRIGG), which ensures the security and confidentiality of the HHR and HHRA and considers issues of data integration or data sharing. Linked data contained within HHRA includes:

- **Primary care data:**
  - Coded clinical entries made during routine patient care
- **Secondary care data:**
  - Radiology and pathology data
  - Inpatient, outpatient and emergency department (ED) secondary uses service (SUS) data.

Primary care coding within the HHRA uses Read codes, which are the standard clinical terminology system used in primary care in the UK. The Read code system comprises a detailed hierarchy of codes describing diagnoses, symptoms, signs, processes of care, investigations, administrative items, procedures and medication; 90% of UK practices use Read version 2 (V2) codes, which were used for the purpose of data capture for this study.

The HHRA allows full access to coded hospital activity data from SUS, thereby allowing quantification of respiratory-related hospital-based health care use to a much greater extent than is possible through databases accessing primary care data alone. The hospital data are the same as that used in contracting, accounting and NHS reporting processes, so high quality is assured.

The secondary care information used was Payment by Results data drawn from the SUS, which involves many quality control/validation processes applied at national level. For primary care information, basic data cleaning had already been performed within the NHS environment before data were imported into the HHR and HHRA. For this study, data cleaning involved excluding from the data set duplicated records and assessments with null and zero values and values deemed incompatible with life.