Leisure time physical activity, sedentary behavior and risk of cardiovascular disease and mortality among US Veterans

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Abstract
Background: Cardiovascular Disease (CVD) remains the leading cause of death in the US. Although Physical Activity (PA) has been inversely associated with the risk of CVD, few studies have examined whether sedentary behaviors modify such association. Our goal was to examine associations of leisure time PA with risk of CVD and mortality and the role of sedentary behavior as potential effect modifier among US veterans.

Methods: We analyzed self-reported data on leisure time PA, television watching, and time spent on the computer among 438,364 participants of the Veterans Affairs Million Veteran Program from 2011 to 2018. We calculated metabolic equivalent of task-hours per week (MET-h/week) for each person and used electronic health record data to ascertain CVD.

Results: Mean age was 64.6 ± 12.6 years and 92% were men. During a mean follow up of 3.3 years, we observed 22,942 new cases of CVD and 48,325 deaths. There was an inverse relation of leisure time PA with CVD and total mortality [HR: 0.96 (0.95-0.97) and 0.91 (0.90-0.92) per 2 MET-h/week increment for CVD and total mortality, respectively]. The associations of PA with both incident CVD and mortality were stronger in participants who spent more time watching television or on computer (all p values for interaction < 0.01). No interaction of PA with time spent on video game was observed (p>0.05).

Conclusions: Leisure time PA is inversely associated with risk of CVD and mortality among US veterans and such relations were stronger in participants who spent more time watching television or on computer.

Introduction
Despite a noticeable decline in deaths from cardiovascular diseases over the past decades, Atherosclerotic Cardiovascular Disease (ASCVD) remains one of the leading causes of death in the United States and is associated with high costs [1]. While modifiable lifestyle factors including Physical Activity (PA) have been reported to reduce the risk of ASCVD in the general population [2-6], limited data are available on the association of leisure time PA and incidence of ASCVD and mortality among US veterans. Both the 2018 Physical Activity Guidelines for Americans [7] and the 2019 CVD Primary Prevention Clinical Practice Guidelines [8] recommend ≥ 150 minutes per week of moderate intensity such as brisk walking (equivalent of 7.5 MET-h per week) or ≥ 75 minutes per week of vigorous-intensity activity (i.e., shoveling snow) for adults for health benefits. Unfortunately, current data suggest that only one in five American adults and less than 25% of US veterans meet these recommended goals for PA [1]. It is less clear whether PA levels below the recommended amount, which are highly
prevalent in older adults, are associated with lower ASCVD or mortality risk among US veterans. Furthermore, sedentary behavior (i.e., hours spent watching television) is highly prevalent in the US with adults spending on average 10.5 hours per day connected on media including television and computer [1]. Sedentary behavior has been positively associated with type 2 diabetes [9], ASCVD [10], and mortality [9,11]. However, it is unclear whether people who spend more time in sedentary behaviors benefit the most from PA compared to those in the lowest categories of sedentary behaviors. Thus, the current project sought to prospectively test the (i) primary hypothesis that leisure time PA is inversely associated with incidence of ASCVD and mortality and (ii) secondary hypothesis that sedentary behavior modifies the association of leisure time PA with incident ASCVD and/or mortality among participants of the Veterans Affairs (VA) Million Veteran Program (MVP).

Methods

Population

MVP is an ongoing prospective cohort study and large biorepository designed to study genetic determinants of chronic diseases among US veterans who use Veterans Health Administration (VHA) services. All veterans signed informed consent and the VA Central Institutional Review Board approved the study protocol in 2010. Details on design and methodology of MVP have been published elsewhere [12]. As of January 19, 2019, 702,740 veterans have been enrolled. Among enrollees, 475,118 participants provided self-report data through the Baseline and/or Lifestyle Survey with questions about modifiable lifestyle factors including leisure time PA. A total of 470,251 answered survey questions on leisure time PA and after exclusion of 31,344 participants with no Electronic Health Record (EHR) follow up data and exclusion of 542 participants with missing age, we used information on 438,364 participants for the current analyses. Details on exclusion criteria are presented in the flow chart (supplementary Figure 1).

Assessment of leisure time PA

Item 24.c of the Lifestyle Survey asked the following question about leisure time PA: “During your leisure or free time, how often do you engage in the following levels of activity?” Three levels were specified as follows: “Vigorous (e.g., competitive sports like running, swimming, or high intensity aerobics)”; “Moderate (e.g., low impact aerobics, or golfing without a power cart)” and “Light (e.g., bowling, archery, easy walking, golfing with a power cart, fishing)”. Response options included: “Never” [0/week]; “Once/month or less” [0.25/week]; “Several times/month”[0.625/week]; “Once per week”[1/week]; “Several times/week”[4/week]; and “Daily” [7/week]. On the Baseline Survey, item # 29 asked participants to report their level of vigorous physical activity: “How often do you exercise vigorously enough to work up a sweat?”. Pre-specified answers were: “rarely/ Never”[0/week]; “1–3 times a month”[0.5/week]; “Once a week” [1/week]; “2–4 times a week”[3/week]; “5–6 times a week”[5.5/week]; and “Daily” [7/week]. We use answers to these questions on leisure time PA to compute MET-h/week. For participants with missing data on leisure time PA (item 24c of the Lifestyle Survey; n=144,700) we used data on vigorous physical activity obtained from the Baseline Survey (item 29) to compute MET-h/week. We made the following assumptions: a) we assigned 3, 5, and 9 METs for light, moderate, and vigorous activity, respectively [13]; b) since we did not query about the duration of each type of physical activity, we assumed fifteen minutes (15 minutes or 0.25 hour) duration for each reported frequency; and c) used the mid-point for answers that specified a range (i.e., we assigned a frequency of 5.5 times a week for “5–6 times a week”) [see supplemental Table 1 for conversion of all answers]. For each of the 3 categories (light, moderate, and vigorous) of leisure time PA, we multiplied corresponding MET by duration (0.25 hour), and the obtained result was then multiplied by the reported frequency in weeks. Total MET-h/week was obtained by summing MET-h/week of light, moderate, and vigorous activity. In sensitivity analyses, we repeated the above algorithm using 20 and 30 minutes in duration and obtained similar correlation between MET-h/week and HDL-cholesterol measured closest to the assessment of PA (Rho: 0.106, 0.106, and 0.106 for use of 15, 20, and 30 minutes, respectively). Since the use of 20 or 30 minutes would exaggerate level of PA among light active veterans, we chose to be conservative and err on the side of underestimation rather than overestimation of PA and retained 15 minute–duration for the current analyses. Once MET-h/wk was calculated for each participant, we grouped leisure time PA into quintiles because the association of leisure time PA was steepest within the first quintile among light, moderate, and vigorous activity. In sensitivity analyses, we repeated the above algorithm using 20 and 30 minutes in duration and obtained similar correlation between MET-h/week and HDL-cholesterol measured closest to the assessment of PA (Rho: 0.106, 0.106, and 0.106 for use of 15, 20, and 30 minutes, respectively). Since the use of 20 or 30 minutes would exaggerate level of PA among light active veterans, we chose to be conservative and err on the side of underestimation rather than overestimation of PA and retained 15 minute–duration for the current analyses. Once MET-h/wk was calculated for each participant, we grouped leisure time PA into quintiles because we did not assume a nonlinear relation between exposure and outcome.

Assessment of sedentary behaviors

Item #30 of the Lifestyle Survey asked participants to provide information about sedentary behaviors: “During the PAST MONTH, on average, how many hours per week did you spend: Watching TV, Video, or DVD; Using a computer; Playing video games; Talking on a cell phone”. Pre-specified answers were: “0, 1, 2–5, 6–10, 11–20, 21–40, 41–60, 61–90, over 90 hours”. We used the mid-point for each category and the floor of the open–ended category for analyses. Since we could not assume that people are immobile while talking on the phone, we chose to focus primarily on “watching TV, video, or DVD” , time “Using a computer” and “Playing video games” as a surrogate of sedentary behaviors.

Assessment of CVD and mortality

CVD included non–fatal myocardial infarction using ICD–9 Codes 410–411, 413–414 and ICD–10 codes I20 – I25 (excluding I25.2); coronary deaths (ICD10 I20–I25); coronary angioplasty or revascularization (CPT Codes 33510–33536, 9292x, 9293x, 9294x, 92973, 92974, and 92975; ICD–9 Procedure codes 36.x and 00.66). Fatal and non–fatal stroke (both ischemic and hemorrhagic stroke): ICD–9 codes 430–431.x, 433–434.x, 436.x, 437.x and 437.6x and ICD–10 codes 160–61.x, 161.x,165.x, 166.x, 167.x, 167.6x, 167.8x ; mortality was ascertained using the National Death Index (NDI) [14]. The validity of using ICD codes for the diagnosis of cardiovascular disease among veterans has been previously published [15]. Furthermore, the use of high throughput methods for phenotyping in the VA has been previously published [16,17].

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weights were estimated using Generalized Boosted Models (GBM) [20] from the weightit [21] and gbm [22] R packages and the following set of variables: age, sex, race, education, BMI, smoking, alcohol intake, and DASH score. For all-cause mortality we additionally adjusted for prevalent CVD at baseline. We considered prevalent diabetes, atrial fibrillation, heart failure, and hypertension as potential intermediate factors in the causal path of PA and CVD and therefore did not control for them. Furthermore, since PA is positively associated with HDL-cholesterol, we did not control for lipids or treatment for dyslipidemia. Stabilized inverse probability of treatment weights (sIPTW) were calculated for each patient by converting the propensity scores. After checking for proportionality of hazards there was no violation and Cox proportional hazards models were used to estimate the crude and inverse probability weighted hazard ratios with 95% confidence intervals. In addition to our main analyses, we evaluated the shape of the PA-outcome relation using restricted cubic splines [23] with knots placed at 25th, 50th, and 75th percentile of continuous MET-h/wk. We examined effect modification by sedentary behavior by conducting stratified analyses using

### Table 1: Baseline characteristics of 438,364 participants of the Million Veteran Program by quintiles of leisure time physical activity.

| Characteristics          | Q1 [0] (n=91,888) | Q2 [0.1-1.24] (n=91,234) | Q3 [1.25-3.80] (n=79,405) | Q4 [3.81-6.60] (n=88,858) | Q5 [6.61-30] (n=86,979) |
|--------------------------|-------------------|--------------------------|--------------------------|--------------------------|--------------------------|
| Age (y)*                 | 65.94 ± 12.3      | 65.07 ± 12.1             | 64.83 ± 12.4             | 64.31 ± 12.9             | 62.91 ± 13              |
| Male Sex (%)             | 84674 (92.1 %)    | 84003 (92.1 %)            | 73051 (92 %)              | 81063 (91.2 %)           | 79121 (91 %)            |
| Race (%)                 |                   |                          |                          |                          |                          |
| White                    | 69805 (76 %)      | 73829 (80.9 %)            | 64853 (81.7 %)            | 71565 (80.5 %)           | 72311 (83.1 %)          |
| Black                    | 17405 (18.9 %)    | 13001 (14.3 %)            | 10477 (13.2 %)            | 12397 (14 %)             | 97712 (11.2 %)          |
| Education (%)*           |                   |                          |                          |                          |                          |
| < High School            | 6191 (7.3 %)      | 3003 (3.7 %)              | 2029 (2.8 %)              | 2293 (2.8 %)             | 1672 (2.1 %)            |
| ≥ High School            | 78314 (92.7 %)    | 78746 (96.3 %)            | 70348 (97.2 %)            | 80084 (97.2 %)           | 78172 (97.9 %)          |
| BMI (kg/m2)*             | 30.15 ± 6.5       | 30.12 ± 5.9               | 29.52 ± 5.4               | 29.06 ± 5.3              | 28.36 ± 4.9             |
| Smoking Status (%)*      |                   |                          |                          |                          |                          |
| Never                    | 9611 (23.2 %)     | 15889 (23.8 %)            | 15279 (27 %)              | 14320 (27.6 %)           | 22032 (33.9 %)          |
| Former                   | 22795 (51.8 %)    | 36908 (55.3 %)            | 31715 (56.1 %)            | 29171 (56.2 %)           | 34885 (53.7 %)          |
| Current                  | 9063 (21.9 %)     | 13966 (20.9 %)            | 9515 (16.8 %)             | 8449 (16.3 %)            | 8008 (12.3 %)           |
| Drinker (%)*             |                   |                          |                          |                          |                          |
| Never                    | 4834 (11.6 %)     | 5109 (7.7 %)              | 4059 (7.2 %)              | 3979 (7.7 %)             | 4234 (6.5 %)            |
| Former                   | 21496 (51.8 %)    | 28687 (43 %)              | 21169 (37.5 %)            | 19589 (37.7 %)           | 20167 (31.1 %)          |
| Current                  | 15165 (36.5 %)    | 32976 (49.4 %)            | 31297 (55.4 %)            | 28385 (54.6 %)           | 40537 (62.4 %)          |
| Diabetes (%)             | 33115 (36 %)      | 29076 (31.9 %)            | 22158 (27.9 %)            | 22570 (25.4 %)           | 16524 (19 %)            |
| Hypertension (%)         | 70461 (76.7 %)    | 66125 (72.5 %)            | 54339 (68.4 %)            | 58585 (65.9 %)           | 50555 (58.1 %)          |
| Dyslipidemia (%)         | 67309 (73.3 %)    | 66169 (72.5 %)            | 55734 (70.2 %)            | 60543 (68.1 %)           | 54939 (63.2 %)          |
| Atrial fibrillation (%)  | 11396 (12.4 %)    | 9257 (10.1 %)             | 7292 (9.2 %)              | 7644 (8.6 %)             | 5992 (6.9 %)            |
| Heart failure (%)        | 11933 (13 %)      | 7958 (8.7 %)              | 5295 (6.7 %)              | 5580 (6.3 %)             | 3508 (4 %)              |
| DASH score*              | 19.49 ± 4.9       | 20.06 ± 4.8               | 21.16 ± 4.8               | 21.66 ± 4.9              | 22.93 ± 4.9             |
| Statin use (%)           | 60541 (65.9 %)    | 57362 (62.9 %)            | 46677 (58.8 %)            | 50034 (56.3 %)           | 43111 (49.6 %)          |

**Important covariates**

We collected self-reported information on age, sex, race, education, Body Mass Index (BMI), alcohol consumption, and smoking through the Baseline Survey. Diet was assessed in MVP using the Willett semi-quantitative food frequency questionnaire, with validity and reproducibility reported previously in other cohorts [18]. We constructed a modified Dietary Approach to Stop Hypertension (DASH) (without dietary sodium since MVP did not have nutrients at the time of current analyses) to characterize overall dietary quality [3]. Prevalent CVD and other comorbidities were derived through the VHA EHR system, Corporate Data Warehouse (CDW) [19], using ICD-9 and ICD-10 codes.

**Statistical analysis**

We censored participants at first occurrence of CVD, death, or last recorded visit and calculated the crude incidence rate by dividing the number of outcomes by corresponding person-time. Given that baseline characteristics among the quintiles of leisure time PA were unbalanced (Table 1), we used multinomial propensity score weighting methods and employed an Average Treatment Effect (ATE) weighting strategy. Propensity score weights were estimated using Generalized Boosted Models (GBM) [20] from the weightit [21] and gbm [22] R packages and the following set of variables: age, sex, race, education, BMI, smoking, alcohol intake, and DASH score. For all-cause mortality we additionally adjusted for prevalent CVD at baseline. We considered prevalent diabetes, atrial fibrillation, heart failure, and hypertension as potential intermediate factors in the causal path of PA and CVD and therefore did not control for them. Furthermore, since PA is positively associated with HDL-cholesterol, we did not control for lipids or treatment for dyslipidemia.

Stabilized inverse probability of treatment weights (sIPTW) were calculated for each patient by converting the propensity scores. After checking for proportionality of hazards there was no violation and Cox proportional hazards models were used to estimate the crude and inverse probability weighted hazard ratios with 95% confidence intervals.

In addition to our main analyses, we evaluated the shape of the PA-outcome relation using restricted cubic splines [23] with knots placed at 25th, 50th, and 75th percentile of continuous MET-h/wk. We examined effect modification by sedentary behavior by conducting stratified analyses using

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tertiles of each of the three sedentary behaviors (time spent on television watching, computer, and video games) and utilizing the product term of leisure time PA and time spent on each sedentary behavior in an sIPTW Cox proportional hazards model to obtain p value for interaction. In sensitivity analysis, we excluded subjects with follow up time <1 year. Propensity score modeling was conducted in R 4.0.2, all other analyses were performed on SAS Enterprise Guide 7.1. An alpha level at 0.05 was used.

Results

Among 438,364 veterans analyzed, mean age was 64.6 ± 12.6; 92% were men; 80.4% white; and 14.4% black. Median leisure time PA was 2.25 MET·h/week and 19% met the federal guidelines recommending at least 7.5 MET·h/week. Frequent leisure time PA was associated with younger age, white race, higher educational attainment, lower body mass index and prevalence of hypertension and diabetes as expected (Table 1). We applied GBM to obtain propensity score weights in order to balance the distribution of baseline characteristics. After adjustment, the absolute standardized mean difference for every covariate was < 0.1 indicating good balance (Supplemental Figure 2). Median time spent on sedentary behavior was 15.5 h/week for television watching, 3.5 h/week for computer, and 0 h/week for video games. During a mean follow up of 3.3 years (range: 0.003 to 7.7 years), 22,942 new cases of CVD and 48,325 deaths occurred. Leisure time PA was inversely associated with risk of CVD with multivariable adjusted hazard ratios (95% CI) of 1.00 (ref), 0.87 (0.82, 0.92), 0.77 (0.73, 0.82), 0.77 (0.73, 0.82) and 0.69 (0.65, 0.73) from the lowest to the highest quintile of PA, using stabilized inverse probability of treatment weights, Table 2. Similarly, we observed an inverse association of leisure time PA with risk of CAD and both ischemic but not hemorrhagic strokes (Supplemental Tables 2–4). Using restricted cubic spline, we found evidence of an inverse and linear relation between leisure time PA and risk of CVD (p non-linear trend <0.0001, Figure 1). Furthermore, leisure time PA was inversely associated with younger age, white race, higher educational attainment, lower body mass index and prevalence of hypertension and diabetes as expected (Table 1). We applied GBM to obtain propensity score weights in order to balance the distribution of baseline characteristics. After adjustment, the absolute standardized mean difference for every covariate was < 0.1 indicating good balance (Supplemental Figure 2). Median time spent on sedentary behavior was 15.5 h/week for television watching, 3.5 h/week for computer, and 0 h/week for video games. During a mean follow up of 3.3 years (range: 0.003 to 7.7 years), 22,942 new cases of CVD and 48,325 deaths occurred. Leisure time PA was inversely associated with risk of CVD with multivariable adjusted hazard ratios (95% CI) of 1.00 (ref), 0.87 (0.82, 0.92), 0.77 (0.73, 0.82), 0.77 (0.73, 0.82) and 0.69 (0.65, 0.73) from the lowest to the highest quintile of PA, using stabilized inverse probability of treatment weights, Table 2. Similarly, we observed an inverse association of leisure time PA with risk of CAD and both ischemic but not hemorrhagic strokes (Supplemental Tables 2–4). Using restricted cubic spline, we found evidence of an inverse and linear relation between leisure time PA and risk of CVD (p non-linear trend <0.0001, Figure 1). Furthermore, leisure time PA was inversely associated with risk of total mortality (Table 3) as well as CVD and cancer mortality (Supplemental Tables 5,6).

The inverse associations of PA with incident CVD or mortality were stronger in the highest tertile of television watching and computer work (all p interaction <0.05, Figure 2). Time spent on video games did not modify PA-CVD or PA-mortality relation (all p for interaction >0.05), data not shown. In sensitivity analyses excluding subjects with follow up time below 1 year did not alter the main results (p linear trend <0.0001).

Discussion

Main findings

In this large and well-characterized cohort of US veterans, we observed an inverse association of leisure time PA with incidence of CVD (CAD and ischemic stroke) as well as mortality (including CVD and cancer mortality) after adjustment for potential confounding factors. No meaningful association

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**Table 2.** Hazard ratios (95% CI) for CVD by leisure time physical activity in the Million Veteran Program (N=315,119).

| Quintiles of leisure time activity (MET-h/week) | Cases/N | Crude incidence (/1000PY) | Crude | sIPTW¹ |
|-----------------------------------------------|---------|---------------------------|-------|--------|
| Q1 (0)                                        | 5,628/58,434 | 30.9 | 1.00 (ref) | 1.00 (ref) |
| Q2 (0.1-1.24)                                  | 4,992/63,262 | 25.1 | 0.81 (0.78, 0.84) | 0.87 (0.82, 0.92) |
| Q3 (1.25-3.80)                                 | 4,022/57,932 | 21.9 | 0.71 (0.68, 0.74) | 0.77 (0.73, 0.82) |
| Q4 (3.81-6.60)                                 | 4,427/66,053 | 21.1 | 0.68 (0.66, 0.71) | 0.77 (0.73, 0.82) |
| Q5 (6.61-30)                                   | 3,873/69,438 | 17.5 | 0.57 (0.54, 0.59) | 0.69 (0.65, 0.73) |
| p linear trend                                 | <0.0001 | <0.0001 |

¹Adjusted for age, sex, race, education, body mass index, smoking, alcohol intake, and DASH score

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**Table 3.** Hazard ratios (95% CI) for total mortality by leisure time physical activity in the Million Veteran Program (N=438,364).

| Quintiles of leisure time activity (MET-h/week) | Cases/N | Crude incidence (/1000PY) | Crude | sIPTW¹ |
|------------------------------------------------|---------|---------------------------|-------|--------|
| Q1 (0)                                        | 16,696/91,888 | 55.6 | 1.00 (ref) | 1.00 (ref) |
| Q2 (0.1-1.24)                                  | 10,677/91,234 | 35.3 | 0.65 (0.63, 0.67) | 0.73 (0.71, 0.76) |
| Q3 (1.25-3.80)                                 | 7,585/79,405 | 28.6 | 0.52 (0.51, 0.54) | 0.61 (0.59, 0.64) |
| Q4 (3.81-6.60)                                 | 7,962/88,858 | 26.8 | 0.49 (0.47, 0.50) | 0.59 (0.57, 0.62) |
| Q5 (6.61-30)                                   | 5,405/86,979 | 18.6 | 0.34 (0.33, 0.35) | 0.48 (0.46, 0.5) |
| p linear trend                                 | <0.0001 | <0.0001 |

¹Adjusted for age, sex, race, education, body mass index, smoking, alcohol intake, DASH score and CVD at baseline.

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of leisure time PA and incidence of hemorrhagic stroke was observed in this cohort. Furthermore, the relations of leisure time PA with CVD and mortality were stronger in participants who spend more time watching television and working on computer. We did not find evidence of interaction between leisure time PA and time spent playing video games and risk of CVD or mortality.

Leisure time PA and risk of CVD and mortality

While physical inactivity has been ranked fourth among the leading risk factors for mortality worldwide [25], current evidence lends support to beneficial effects of leisure time PA on cardiometabolic risk [1]. In a meta-analyses of prospective cohorts [26], high level of leisure time PA was associated with a 39% lower risk of CVD (95% CI:25% to 48%) compared to low level of leisure time PA; corresponding reduction for moderate level of leisure time PA was 23% (95% CI:10% to 33%). Furthermore, there was an inverse and graded relation between leisure time PA with incidence of coronary artery disease and stroke in the same meta-analysis [26]. These results are consistent with the reported findings from MVP.

Our findings of inverse relation of leisure time PA with total and CVD mortality are also consistent with a meta-analysis of 44 prospective cohorts reporting a linear and inverse association of leisure-time PA with risk of cardiovascular mortality irrespective of age, sex, and prevalent CVD [27]. Our findings extend our current knowledge from existing data by also reporting inverse relation between leisure time PA with cancer mortality.

Interaction of sedentary behavior with leisure time PA on risk of CVD/mortality

Limited studies have examined the interaction of sedentary behavior with leisure time PA on the incidence of CVD and/or mortality. In the 45 and Up Study, sitting time was positively associated with total and CVD mortality in participants that did not exercise or those that did not meet recommendations for PA (at least 150 minutes of moderate to vigorous PA per week) after a median follow up of 8.9 years (p interaction sitting time x PA <0.001) [28]. In contrast, sitting time was not associated with incidence of total or CVD death in participants who met PA guidelines [28]. In the Danish Health Examination Survey (2007–2008), the positive association of sitting time with all-cause mortality was stronger in people that were physically inactive than those who exercised regularly (p interaction <0.05) after 5.4 years of mean follow up; however, no interaction of PA with sitting time was observed for CHD [29]. In contrast, data from the prospective Women’s Health Initiative Observational Study [30] showed a positive and linear

Figure 2: Forest plots of hazard ratio (95% CI) of PA with incident CVD and mortality stratified by (i) computer time (panels A and B) and (ii) TV watching (panels C and D), respectively.
relation of sitting time with incidence of CVD, irrespective of PA after a median follow up of 12.2 years (p for interaction between sitting time and PA 0.94).

Biologic mechanisms

PA reduces the risk of CVD and mortality through its beneficial effects on cardiometabolic risk factors. Several randomized controlled trials have reported that an intervention with PA led to increased HDL-cholesterol, improvement of insulin sensitivity and resistance, and reduction in blood pressure and measures of adiposity [31-33]. Other randomized clinical trials have demonstrated beneficial effects of PA on inflammatory cytokines [34, 35], blood pressure [36], lipids [37], beta cell function [38], and oral glucose tolerance [39].

Limitations and strengths of the study

Our study has some limitations. First, we relied on self-reported PA and sedentary behaviors for current analyses and cannot exclude exposure misclassification in the data. Second, despite multivariable adjustment, unmeasured and/or residual confounding might still explain partially or completely observed association, given the observational design of our study. Third, it is possible that we missed some CVD events that occurred outside the VHA; however, those events would be minimal given the fact that medical records on care provided outside VA are sought after and captured in EHR. Fourth, we did not have information on duration of each bout of PA for accurate assessment. Fifth, MVP participants may not represent the entire US veteran population nor the overall US population for generalization of our findings. Nonetheless, this study has several strengths including a large sample size and adequate number of events to allow subgroup analyses; availability of data on major confounding factors for multivariable adjustment; complete ascertainment of mortality in the VHA, and the availability of adequate data on hemorrhagic stroke for subtype analyses.

Conclusion

Our data showed an inverse association of leisure time PA with risk of CVD and mortality with stronger effect size in participants with a higher propensity of adhering to sedentary behaviors among US veterans. This finding underscores the importance of following PA recommendations as a cost-effective strategy to reduce the burden of CVD. Future studies are needed to explore biologic mechanisms underlying observed interaction of PA with sedentary behaviors in the general population.

Declarations

Ethics approval and consent to participate: Each participants signed informed consent and the study was approved by the Central VA Institutional Review Board, Washington DC.

Availability of data and materials

The datasets generated and/or analyzed during the current study are not publicly available in order to comply with the Department of Veterans Affairs Office of Research and Development policies designed to protect US Veterans’ information.

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Authors’ contribution

Luc Djousse: Designed the study, curated physical activity variable, directed statistical analyses, and drafted the manuscript.

Petra Schubert: Assisted with data curation and completed statistical analyses

Yuk-Lam Ho: Assisted with data curation, data analysis, quality control check of the SAS code

Stacey B. Whitbourne: Collected data and participated in data curation and analysis.

Kelly Cho: Contributed to data collection, supervision of analytical team, and curation of phenotypes

J. Michael Gaziano: Serves as one of the Principal Investigators of MVP, designed MVP study, secured funding, and supervised the team.

All co-authors contributed to data interpretation and approved the final version of the manuscript.

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(Supplemental Tables and Figures)

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