A Proof of the Vieille result using a kind of discount factor

Robert Samuel Simon
Abstract:

We give an alternative proof that every two-person non-zero-sum absorbing positive recursive stochastic game with finitely many states has approximate equilibria, a result proven by Nicolas Vieille. Our proof uses a state specific discount factor which is similar to the conventional discount factor only when there is only one non-absorbing state. Additionally we show that if the players engage in time homogeneous Markovian behavior relative to some finite state space of size $n$ then for the existence of an $\epsilon$-equilibrium it suffices that one-stage deviation brings no more than an $\epsilon^3/(nM)$ gain to a player, where $M$ is a bound on the maximal difference between any two payoffs.

Key words: Stochastic Games, Markov Chains

This paper would not have been possible without the very generous help of Eilon Solan; furthermore some of the paper was written by him. This research was supported by the Center for Rationality and Interactive Decision Theory (Jerusalem), the Department of Mathematics of the Hebrew University (Jerusalem), the Edmund Landau Center for Research in Mathematical Analysis (Jerusalem) sponsored by the Minerva Foundation (Germany), the German Science Foundation (Deutsche Forschungsgemeinschaft), and the Institute of Mathematical Stochastics (Goettingen).
1 Introduction

A two-player stochastic game is played in stages. At every stage the game is in some state of the world. Both players are informed of the whole history, including the current state, and based on this information they choose simultaneously a pair of actions. The current state and the pair of actions chosen determine both a stage payoff for each of the players and a probability distribution according to which a new state is chosen.

For any $\epsilon \geq 0$, an $\epsilon$-equilibrium in a game is a set of strategies, one for each player, such that no player can gain in payoff by more than $\epsilon$ by choosing a different strategy, given that all the other players do not change their strategies. A game has approximate equilibria if for every positive $\epsilon > 0$ it has an $\epsilon$-equilibrium. The value of a zero-sum game, should one exist, is the unique cluster point of the $\epsilon$-equilibrium expected payoffs (for the first player) as $\epsilon$ goes to zero. The un-discounted payoff of a player in a stochastic game with infinitely many stages, when defined, is a limit as the number of stages goes to infinity of the average summed over the stages of the player’s expected payoffs. Unless specified, the payoffs of a stochastic game are undiscounted.

Shapley (1953) presented the model of stochastic games, and proved that a discounted zero-sum games always have a value obtainable with stationary optimal strategies. This result was generalized for equilibria in $n$-player non-zero-sum discounted games by Fink (1964).

An absorbing state is such that the play never leaves this state once it is reached. Kohlberg (1974) proved that every two-player zero-sum stochastic game with only one non-absorbing state has a value. Based on the work of Bewley and Kohlberg (1976), Mertens and Neyman (1981) generalized this result, and proved that every zero-sum stochastic game has a value.

A stochastic game is recursive if the stage payoff at all non-absorbing states is zero, no matter what the players do. A recursive stochastic game is positive recursive if there is a player who receives at all absorbing states only positive payoffs. A positive recursive stochastic game is absorbing if the player who receives these positive payoffs can force the play toward absorption.

Existence of approximate equilibria in two-player non-zero-sum stochastic games with only one non-absorbing state was proven by Thuijsman and Vrieze (1989). In their proof Thuijsman and Vrieze considered a sequence of stationary equilibria of the discounted game as the discount factor tends to
1, and they constructed different types of $\epsilon$-equilibrium strategies according to various properties of the sequence.

Vieille (2000a) showed that for approximate equilibria to exist in every two-player non-zero-sum stochastic game with finitely many states it is sufficient to prove this for the sub-class of absorbing positive recursive games. Furthermore Vieille (2000b, 2000c) proved that indeed all games in this sub-class have approximate equilibria.

In the present paper, we provide an alternative proof of the Vieille result for absorbing positive recursive games. The primarily difference between our proof and Vieille’s lies in the use of a kind of discount factor rather than Vieille’s undiscounted evaluation. This discount factor is state specific and is similar to the conventional discount factor only when there is only one non-absorbing state. We were inspired by the Thuijsman and Vrieze article and their confidence that their ideas could deliver the same result for finitely many states. Our goal was to confirm their optimism by demonstrating the great versatility of the discounting concept.

In positive recursive games, discount factors for the player receiving positive absorbing payoffs persuade him to make moves that push the game toward absorption. Let us call this player the second player. The serious problem with generalizing the Thuijsman and Vrieze approach directly is that the usual discounted evaluation does not discriminate between the time spent at the state at which a decision is made and the other states that might follow this decision. As long as the second player at a given state chooses between two moves that do not involve returning to that state, his evaluation of those moves in an appropriate discounted game should be based upon his undiscounted evaluation. Play that never returns to this state before absorption but visits other states arbitrarily many times receives no discount whereas play that re-visits the initial state $n$ times receives a $(1-\delta)^n$ discount, regardless of its visits to other states.

We see no way to generalize our proof to three player games (and it appears highly unlikely). On the other hand, we can not dismiss the possibility; (see also Solan, 1999, where discounted evaluations were used to understand some three player undiscounted stochastic games). If the compactification of a strategy space creates discontinuities in the undiscounted payoffs a discounted evaluation may handle the points of discontinuity successfully. A false impression that discounting is useless to understanding the undiscounted game may result from a lack of knowledge of how to turn off
the discount where one is sufficiently far from the points of discontinuity. As we will see below, knowing when to turn off the discount is central to our approach.

The secondary difference between our proof and Vieille’s is that the mathematics we use is entirely elementary. No deep theorems of mathematics are required; for example, there is no use of the theory of semi-algebraic functions. What we need from the theory of Markov chains is very elementary and proved entirely in this paper. Due to our discounting approach we work with taboo probabilities rather than the directed graphs perspective of Freidlin and Wentzell, (1984).

The only theorem we quote instead of proving is Doob’s submartingale inequality, a generalization of Kolmogorov’s inequality and also an easy theorem to prove. Applying the inequality, we show that if the players engage in time homogeneous Markovian behavior relative to some finite state space of size \( n \) then for the existence of an \( \epsilon \)-equilibrium it suffices that one-stage deviation brings no more than an \( \epsilon^3/nM \) gain to a player, where \( M \) is a bound on the maximal difference between any two payoffs.

Countably many states

We developed our unorthodox approach to stochastic games with the hope that it would deliver approximate equilibrium existence for all two-person non-zero-sum stochastic games with countably many states. We have failed in this attempt.

The main problem is that our approach (and that of Vieille) rests ultimately on the pigeon-hole principle. If the expected number of visits to every non-absorbing state is finite then with probability one an absorbing state is reached. This does not hold if there are infinitely many non-absorbing states.

In general, what is the difficulty in proving approximate equilibrium existence for non-zero-sum two-person stochastic games with countably many states? Several important positive results need to be mentioned. Maitra and Sudderth (1991) proved that all zero-sum stochastic games with countably many states have values. In a game of perfect information, the players take turns making their moves and each player knows the previous moves of the other players; the classic example is that of chess. A Blackwell game is identical in transition structure to a stochastic game, but the payoffs are determined by a function Borel measurable with respect to the histories of play. Martin (1975) proved that all zero-sum Blackwell games of perfect
information have values, and Mertens and Neyman (in Mertens 1987) ex-
tended Martin’s result to non-zero-sum games with finitely many players. 
Using his result for games of perfect information, Martin (1998) proved that 
all zero-sum Blackwell games have values.

The differences between non-zero-sum stochastic games (with simultane-
ous moves) and either non-zero-sum Blackwell games of perfect information 
or zero-sum Blackwell games with simultaneous moves are formidable. The 
probability of absorption at a stage in a stochastic game can be also a min-
imal bound on that stage’s deviation from pure equilibria; (for example see 
the “Big Match” in Blackwell and Ferguson, 1968). With the \( \epsilon \)-equilibria 
of many games, including the absorbing positive recursive variety, while ab-
sorption must become a near certainty the cumulative opportunity to exploit 
deviations must not exceed \( \epsilon \). Therefore one needs that stage for stage ap-
proximate equilibria can translate to cumulative approximate equilibria. In 
zero-sum games this is not so problematic because the gains to one player 
from deviation equal the losses to the other player. But with two-person 
non-zero-sum games, one must consider functions with values in \( \mathbb{R}^2 \); the po-
tential independence of the two values and need for a cooperative solution 
frustrate attempts to generalize the approaches that were successful with zero-
sum games. On the other hand if the moves are made simultaneously how 
does one know the other player is adhering to a cooperative agreement? So 
far the main answer has been to request from each player Markovian behav-
ior, accompanied by statistical testing and punishment by the other player 
in the event of significant statistical deviation. With this approach, it is nec-
essary that the probability that an honest player will be punished unjustly 
can be made arbitrarily small. As we will demonstrate with the following 
proposition and counter-example to a variation on this proposition, such a 
control process is unlikely in general for Markovian behavior that is carried 
out essentially on a countable state space.

If \( S \) is a finite or countable set let \( \Delta(S) \) stand for the space of probability 
distributions on \( S \). A Markov chain is defined by a finite or countable state 
space \( S \) and for every \( s \in S \) and stage \( i \geq 0 \) a probability distribution 
\( p_i^s \in \Delta(S) \) governing the distribution on the states at the \( i + 1 \)st stage, 
given that \( s \) is the state on the \( i \)th stage. It is time homogeneous if \( p_i^s \) is 
indendent of the \( i \).

**Proposition 4.2:** Let \( X \) be a finite space. For every \( x \in X \) let \( Y_x \) be a
finite space, with $Y := \cup_{x \in X} Y_x$. (In the context of stochastic games, $X$ will be the state space and $Y_x$ will be the set of moves that a player has at the state $x \in X$.) There are probability transitions $(p^x \in \Delta(Y_x) \mid x \in X)$ from $X$ to $Y$ and there are probability transitions $(p^y \in \Delta(X) \mid y \in Y)$ from $Y$ to $X$, so that for every starting point $x_0 \in X$ a time homogeneous Markov chain on $X \cup Y$ is defined. On the even stages $i = 0, 2, 4, \ldots$ the process is in $X$ and on the odd stages the process is in $Y$. Let there be an evaluation function $v : X \cup Y \to \mathbb{R}$ that is harmonic with respect to the transitions (meaning that a martingale is formed). Let $M > 0$ be a uniform bound for the maximal difference between all values of $v$. For every pair $x \in X$ and $y \in Y_x$ such that $y$ is reached from $x$ with positive probability (according to $p^x$) the difference between $v(y)$ and $v(x)$ is no more than $\delta > 0$.

**Conclusion:** If $|X| = n$, $\epsilon < 1/2$, and $\delta \leq \epsilon^3/Mn$ then the probability that there exists an $l$ with $\sum_{i=0,2,\ldots}^l (v(y_{i+1}) - v(x_i)) \geq \epsilon$ does not exceed $\epsilon$.

The complexity of the $Y_x$ play no role in the proof of Proposition 4.2, and therefore it could have many generalizations corresponding to variations in the structure of the $Y_x$.

To emphasize the importance of the finite number $|X|$, the following is a counter-example to Proposition 4.2 if we assume that the bound for $\delta$ is independent of the cardinality of $X$. Furthermore, if we consider processes that are not time-homogeneous, it does not help if for every stage the sum over the states of the maximal differences add up to no more than $\delta$.

Consider a random walk on $n + 1$ positions such that at the left end (at position 0) the player receives an absorbing payoff of 0 and on the right end (at position $n$) an absorbing payoff of 1. The space $X$ is the $n + 1$ positions and for every $x \in X$ the two-set $Y_x$ consists of the two directions “left” and “right”. Given any small $\delta > 0$, one can make $n$ large enough so that at every stage the change in expected payoff does not exceed $\delta$. Now reformulate the random walk so that at the $k$th stage of play there is no motion at any $i$ position with $i \neq k \pmod{n - 1}$, but at the $k' = k \pmod{n - 1}$ position there is an equal $1/2$ probability of moving either to the position $k' - 1$ or to $k' + 1$. At each stage the sum over the states of the differences in expected payoffs remains no more than $\delta$, and yet we are no closer to satisfying the conclusion of the proposition. (With $n$ even and starting in the middle position with an expected payoff of $1/2$, for every small positive $\epsilon$ with probability close to $1/2$ there will be motion to a position with an expected payoff of at least...
We expect no proof of approximate equilibrium existence for all non-zero-sum stochastic games with countable state spaces without a radically different approach. If a proof for countably many states can be found, its application to finite state truncations of the countable state game would provide approximate equilibria of the finite state games such that the average number of stages before absorption would not explode with the increase in the finite number of these states. In the proof below for a fixed $\epsilon$ there is no lower bound determined by the number of states on the rate for which an absorbing state is reached. Indeed, because such a proof would imply the existence of yet another alternative proof for finitely many states with dramatic absorption rate properties, we suspect that there is a counter-example. Furthermore, it is possible that the complexities from countably many states involved in a two-player counter-example could be mimicked by the introduction of more players in a stochastic game with finitely many states, yielding a counter-example to approximate equilibria in this context as well.

We suspect that approximate equilibrium existence for a broad class of two-person stochastic games played on countable state spaces must rest on a fundamental assumption: that there is a uniform bound on the number of states possible on any given stage of play. With a finite number of such positions, it is still not clear how appropriate Markovian should be found. Even with only one non-absorbing position, the possible infinite variations, including the number of moves for each player and the order in which similar “types” may appear, make the problem formidable. At least the generalization of Lemma 4.1 to Markov chains that are not time homogeneous will be necessary. Another reason to present our alternative proof of the Vieille result is the hope that it will be relevant to this case, which we call the case of finitely many positions. If for each non-absorbing position one could find an appropriate common identity to an infinite sub-sequence of states occurring in that position, then the pigeon hole principle could be applied successfully. Throughout this paper, we comment on the case of finitely many positions.

**Organization**

To execute our proof efficiently, we will assume that Player One has the ability to send signals to Player Two that are independent of the transitions in the games. The easiest way to formalize this property is to assume that every move of Player One at a non-absorbing state is paired with another move
at the same state that is its identical copy with respect to the transitions. Without this assumption, the proof is formality more involved, less elegant, however essentially equivalent. In the section following the conclusion of the main proof, we prove the result without this signaling assumption.

The argument and the paper are organized as follows.

Section 2 introduces the model of absorbing positive recursive stochastic games and the basic concepts of Markov chains. Additionally we introduce an important concept with regard to the movement between states, called taboo probabilities. A taboo probability is the probability that one moves from an initial state to some set of target states without travelling through some second set of “forbidden” states.

Section 3 gives proofs of all the needed lemmas on Markov chains. The most central lemma is Lemma 3.2; it states that when motions at a multitude of states are removed whose frequencies are only a small fraction of the total motion toward a fixed state then the flow continues toward this fixed state with about the same or greater tendency.

Section 4 contains a proof of Proposition 4.2, which also establishes general sufficient conditions for the existence of approximate equilibria. We create new states from our old states, which we call situations; at most three situations are created from each original state. The method of creating the situations we call polarization, introduced in Section 3. Except for the rare possibility of punishment, our behavior strategies will be stationary on the situations. Section 4 concludes with Theorem 1, a demonstration of sufficient conditions for approximate equilibrium existence in our games.

In Section 5 we introduce the state specific discounted evaluation for the second player. We define the discounted evaluation such that the discounting rates are adjusted for states sufficiently close together, according to a metric determined by the strategies. We select a quantity $\tau$ much smaller than $\epsilon$, and define the discounted evaluation so that moves with more than an $\tau$ probability of non-return to the state are evaluated in an undiscounted way and moves with a $\gamma$ probability of no return with $\gamma < \tau$ are evaluated as if their probability of no return was $\gamma/\tau$. Our choice for $\tau$ is guided by Proposition 4.2.

A serious problem with the state specific discounted evaluation is that the motivations of the second player at one state can be very different from that at another state. Essentially the second player becomes a multitude of players, one for each state. This allows for the second player at some states to
prefer moves that result in too slow a motion toward absorption and therefore also discounted evaluations below the zero-sum value. To avoid this problem, in Section 2 we define a new correspondence, called the “jump” correspondence, based upon stationary strategies optimal in the conventionally discounted game. The use of the jump correspondence by the second player results in fast absorption. The “best-reply” correspondence of the second player is a combination of the jump correspondence with a maximization of the state specific discounted evaluation – when the discounted evaluation is too low, the jump correspondence is activated. For the first player, the undiscounted evaluation is used to define her “best-reply” correspondence. With the “best-reply” correspondences for both player defined, we demonstrate two important properties. Lemma 5.4 shows that at a fixed point the jump correspondence of the second player has only very limited influence on the play. Lemma 5.5 contains the key argument to our entire approach; it is used repeatedly to solve the most difficult problems. It shows that if there is a meaningful discrepancy between the discounted and undiscounted evaluations for the second player then the second player seeks primarily motion with the fastest absorption rate.

The synthesis of the previous sections lies in Section 6. Theorem 2 proves that the conditions of Theorem 1 are always satisfied – implying the existence of approximate equilibria. Here we consider sets such that a significant proportion of all the motion leaving these sets are from Player Two moves with payoffs for Player Two significantly below the set-average payoff. Fixing any such state in a set where such moves take place, we look at what happens when Player One stops playing all moves performed with frequencies small compared to the motion toward this special state. The result, for which Player One is indifferent, involves almost exclusively the use of similar such moves by Player Two such that the players can travel between these moves without the danger that along the way Player Two prefers to provoke punishment over performing one of these moves. Ultimately we show that there is a convex combination of such moves that all yield the same payoff for Player Two and for which Player One is approximately indifferent.

In Section 7 we consider the problem of signaling, as described above; and in Section 8 we conclude in more detail with the problem of countably many states.
2 Preliminaries

2.1 The Model

Let $\mathcal{S}$ be the set of states; $\mathcal{A}$ is the subset of absorbing states and $\mathcal{N} = \mathcal{S} \setminus \mathcal{A}$ is the subset of non-absorbing states.

For every $s \in \mathcal{S}$, $A^s_1$ and $A^s_2$ are the moves (pure actions) of the first and second players, respectively, at the state $s$. Without loss of generality, we assume that $|A^s_i| = 1$ for every $s \in \mathcal{A}$ and $i = 1, 2$. Let $r^1 : \mathcal{A} \to [-1/2, 1/2]$ and $r^2 : \mathcal{A} \to [\omega, 1]$ be the first and second players' evaluations on absorbing states, respectively, with $0 < \omega < 1$. Let $m$ be the maximal number of moves for either player at any non-absorbing state, meaning $m = \max_{s \in \mathcal{N}}(|A^s_1|, |A^s_2|)$.

Let $p(t|s; a, b)$ be the probability of moving from $s$ to $t$ when $a \in A^s_1$ and $b \in A^s_2$ are played. Let $\rho$ be defined by $\rho := \min_{s, t \in \mathcal{S}} p(t|s; a, b) > 0$, the minimal non-zero transition probability. Notice that in the case of finitely many positions one has such a positive quantity for each stage. More relevant, however, would be a sequence $\rho_i$ of positive quantities such that the series $\rho_i$ is divergent but sums toward infinity much slower than any divergent series of positive transition probabilities. Such a series is possible if there is a uniform bound on the number of moves. Additionally the discount factor must be adjusted to this series, (possibly with the discount factor equaling $1 - \delta\rho_i$ if there is only one non-absorbing state).

Let $X := \prod_{s \in \mathcal{N}} \Delta(A^s_1)$ and $Y := \prod_{s \in \mathcal{N}} \Delta(A^s_2)$ be the spaces of stationary strategies of the players, with $X^s := \Delta(A^s_1)$ and $Y^s := \Delta(A^s_2)$. For $a \in A^s_1$, $b \in A^s_2$, $x^s \in X^s$ and $y^s \in Y^s$ we define $p(t|s; a, y^s)$, $p(t|s; x^s, b)$ and $p(t|s; x^s, y^s)$ in the appropriate linear or bi-linear way. For any $s \in \mathcal{N}$, $x^s \in X^s$ and $a \in A^s_1$, the quantity $x^s_a$ will stand for the probability, as determined by $x^s$, that the move $a$ is used. The same applies for $b \in A^s_2$, $y^s \in Y^s$ and $y^s_b$. Define a pair $(x, y) \in X \times Y$ to be absorbing if from every start with probability one an absorbing state is reached.

We will say that two positive quantities $a$ and $b$ are different by no more than a factor of positive $\gamma < 1$ if $a \geq b(1 - \gamma)$ and $b \geq a(1 - \gamma)$. 
2.2 Histories, Strategies, Equilibria

For every stage $i \geq 0$ and $s \in S$ define $H_i^s := \{(s_0, a_0, b_0), (s_1, a_1, b_1), \ldots, (s_{i-1}, a_{i-1}, b_{i-1}), s_i = s \mid \forall 0 \leq k < i a_k \in A_1^{s_k}, b_k \in A_2^{s_k}, p(s_{k+1}|s_k; a_k, b_k) > 0\}$, with $H_0^s = \{s\}$ for all $s \in S$. Define $H^s := \bigcup_{i=1}^{\infty} H_i^s$, $H_i := \bigcup_{s \in S} H_i^s$, $H := \bigcup_{i=0}^{\infty} H_i$, and $\tilde{H} := \{(s_0, a_0, b_0), (s_1, a_1, b_1), \ldots \mid \forall i \geq 0$ the truncation up to $s_i$ belongs to $H_i^s\}$, the set of infinite sequences.

A strategy of Player $j = 1, 2$ is a set of maps $\sigma_j = (\sigma_j^s \mid s \in N)$ with $\sigma_j^s$ a map from $H^s$ to $\Delta(A_j^s)$ for all $s \in N$.

With Blackwell games, a more general class than stochastic games, we assume that a player’s evaluation on $\tilde{H}$ is a function that is measurable with respect to the Borel subsets of $\tilde{H}$, the sigma algebra induced by the subsets of $H_i$ for all $i \geq 0$. In case that a stochastic game is recursive, for every member of $\tilde{H}$ it easy to define an evaluation for both players. Either the infinite sequence reaches an absorbing state and the players receive the corresponding absorbing payoffs, or it never reaches an absorbing state and both players receive a payoff of zero.

For every initial state $s$ and every pair of strategies $\sigma_1, \sigma_2$ for both players a distribution is induced on $\tilde{H}$ in a natural way, resulting in two evaluations $V_j^s(\sigma_1, \sigma_2)$ for Player $j = 1, 2$ of the expected values of the $r_j$ on $\tilde{H}$. An $\epsilon$-equilibrium is a pair $\tilde{\sigma}_1, \tilde{\sigma}_2$ such that for all $s \in S$ and alternative strategies $\bar{\sigma}_1$ and $\bar{\sigma}_2$ it holds that $V_j^s(\bar{\sigma}_1, \bar{\sigma}_2) \leq V_j^s(\tilde{\sigma}_1, \tilde{\sigma}_2) + \epsilon$ and $V_2^s(\tilde{\sigma}_1, \tilde{\sigma}_2) \leq V_2^s(\tilde{\sigma}_1, \tilde{\sigma}_2) + \epsilon$.

With absorbing positive recursive games and positive $\omega$ the lowest Player Two absorbing payoff we get the additional property that there exists an $N > 0$ such that with probability at least $1 - \frac{2\epsilon}{\omega}$ the game has reached an absorbing state before the stage $N$.

2.3 Jump Function

For any positive real number $0 < \alpha < 1$ let $G^\alpha$ be the conventionally defined discounted zero-sum game played against Player Two such that a visit to any state is discounted according to $1 - \alpha$, and let $G^0$ be the corresponding undiscounted zero-sum game. For all positive $\alpha$ we define $c^\alpha : S \rightarrow \mathbb{R}$ to be the min-max value for Player Two in the zero-sum game $G^\alpha$, with $c^\alpha(s) = r^2(s)$ for all $s \in A$. Because the game is positive recursive the $c^\alpha$ are monotonically non-decreasing and due to Mertens and Neyman (1981) the point-wise limit is the undiscounted value of the game $G^0$, though for this
class of games there is an elementary proof. Player Two chooses a stationary optimal strategy of \( G^\alpha \) for an \( \alpha > 0 \) sufficiently small so that \( c^\alpha \) is within \( \epsilon \) of its point-wise limit and at stage \( i \) Player One chooses one of her optimal strategies in the game \( G^{\alpha_i} \) where for every \( i \geq 0 \) \( c^{\alpha_i} \) is within \( \epsilon/2^{i+2} \) of the point-wise limit and \( \alpha_i < \epsilon/2^{i+2} \).

For every \( x \in X \) and positive \( 0 < \alpha < 1 \) define the **jump** function \( j^\alpha_x : N \rightarrow \mathbb{R} \) by

\[
j^\alpha_x(s) = (1 - \alpha) \max_{b \in A^2_s} \sum_{t \in S} p(t|s, x, b) c^\alpha(t)
\]

- the maximal payoff that Player Two can guarantee himself in the \( 1 - \alpha \) discounted game by being punished after the next stage if Player One uses \( x \) at the present stage. If \( s \) is an absorbing state, define \( j^\alpha_x(s) \) to be \( r^2(s) \) for all \( \alpha \). For all states it is clear that \( j^\alpha_x \geq c^\alpha \), with equality when \( x \) is an optimal strategy for Player One in the zero-sum game \( G^\alpha \) played against Player Two.

For every state \( s \in N \) and \( x \in X \) define

\[
J^\alpha_x(s) = \arg\max_{b \in A^2_s} \sum_{t \in S} p(t|s, x, b) c^\alpha(t).
\]

Let \( n(s) \) denote the state following \( s \), in our context a random variable. If \( s \) is not an absorbing state and \( b \in J^\alpha_x(s) \) then \( j^\alpha_x(s) \leq (1 - \alpha) E^x_b j^\alpha_x(n(s)) \), where \( E^x_b \) is the expectation determined by the move \( b \) and the strategy \( x^s \). This makes \( j^\alpha_x \) a sub-martingale.

For \( i = 1, 2 \) and a state \( s \in S \) define \( c_i(s) \) to be the value for Player \( i \) of the zero-sum undiscounted game played against Player \( i \) starting at the state \( s \). For every Player \( i \) and every stationary strategy \( z \) of Player \( k \neq i \) define the **jump** function \( j^i_z : S \rightarrow \mathbb{R} \) by

\[
j^i_z(s) = \max_{b \in A^2_s} \sum_{t \in S} p(t|s, z, b)c_2(t) \quad \text{or} \quad j^1_z(s) = \max_{a \in A^1_s} \sum_{t \in S} p(t|s, a, z)c_1(t)
\]

- the maximal payoff that Player \( i \) can guarantee himself against \( z \) if he is punished on the next stage.

### 2.4 Taboo probabilities

For any time homogeneous Markov chain, a state \( s \), and two disjoint sets \( A \) and \( B \) of states we introduce the “taboo” probability \( P^A(s, B) \) to be the
probability, with a start at the state \( s \), of reaching the set \( B \) before the set \( A \) at any stage following the initial stage at \( s \). With \( t_C := \inf\{n \geq 1 \mid s_n \in C\} \) \( P^A(s, B) \) measures the event that \( t_B < \infty \) and \( t_B < t_A \) conditioned on \( s_0 = s \). If either set is a singleton, we can write its single member instead of the set. If there is ambiguity concerning which state space or which transitions, we identify them with a subscript. In our context of stochastic games and stationary strategies, \( P^A_{x,y}(s, B) \) will be the taboo probability corresponding to the time homogeneous Markov chain generated by \((x, y) \in X \times Y\).

Define a state of a time homogeneous Markov chain to be absorbing if once this state is reached then the motion remains in this state forever. The Markov chain is absorbing if for any start with probability one an absorbing state is reached.

Before moving toward the proof, we must present some basic notions using the taboo probabilities. These quantities will be defined first for time homogeneous Markov chains and then applied to the games.

For any part \( p \) of a transition at a state \( s \) or an alternative transition \( p \) for that state define \( g(p) \) to be the probability that there is no return to \( s \) if \( p \) is used at \( s \) and the transitions remain constant at all other states. If \( p \) was a part of the transition at \( s \) then define \( f_p \) to be the frequency with which \( p \) is used at the state \( s \). For every choice \((x, y) \in X \times Y\) and pair \( a \in A_1^s \) and \( b \in A_2^s \) of moves at the state \( s \in \mathcal{N} \) \( g_{x,y}(a, b) \) is the probability that there is no return to \( s \) given that Player One and Player Two at \( s \) play the actions \( a \) and \( b \), and elsewhere in the future the stationary strategies \((x, y)\). For a move \( b \in A_2^s \) of the second player, define \( g_{x,y}^{b} \) to be \( \sum_{a \in A_1^s} x_a y_b g_{x,y}(a, b) \), and define \( g_{x,y}^a \) for all \( a \in A_1^s \) correspondingly.

Define the absorption rate \( a(s) \) of a state \( s \) to be the probability that after any visit to this state there is no return to this state, meaning that the absorption rate is the expected value of the function \( g \). For the game the absorption rate \( a_{x,y}(s) \) of a state \( s \) is \( \sum_{a \in A_1^s} \sum_{b \in A_2^s} x_a y_b g_{x,y}(a, b) \). Given that \((x, y) \) is absorbing \( a_{x,y}(s) \) would be the taboo probability \( P_{x,y}^{x,y}(s, A) \).

For any part \( p \) of the transition at a state \( s \) define \( \nu(p) \) to be the probability that at the last visit to \( s \) the part \( p \) was used, or equivalently \( \nu(p) = f_p g(p) / a(s) \). We call this the importance of \( p \). For a pair of moves \( a \in A_1^s \) and \( b \in A_2^s \) at \( s \in \mathcal{N} \) and stationary strategies \((x, y)\) the importance \( \nu_{x,y}^{a,b}(a, b) \) is \( x_a y_b g_{x,y}(a, b) / a_{x,y}(s) \). For any move \( a \in A_1^s \) define \( \nu_{x,y}^a \) to be \( \sum_{b \in A_2^s} \nu_{x,y}(a, b) = x_a g_{x,y}^a / a_{x,y}(s) \) and for any move \( b \in A_2^s \) define \( \nu_{x,y}^b \) in the same way.

For any distinct pair \( s, t \) of states define \( \text{esc}(t, s) \) to be the probability of
never reaching \(s\) with a start at \(t\). (esc stands for “escape”.) For the game we have \(g_{x,y}^b = \sum_{t \in S} p(t|s; x, b)esc_{x,y}(t, s)\). (If \((x, y)\) is absorbing, \(esc_{x,y}(t, s)\) is \(P^s_x(t, A)\) and is different from \(P^s_{x,y}(t, A)\), the probability of absorbing before returning to either \(s\) or \(t\).

For distinct states \(s\) and \(t\) let \(\mu(s, t)\) be \(esc(s, t) + esc(t, s)\), and otherwise let \(\mu(s, s) = 0\). \(\mu\) is a metric on the state space. Recognize \(1 - esc(t, s)\) as the probability of moving from \(t\) to \(s\), and for mutually distinct \(u, v, w\) we have \(1 - esc(u, w) \geq (1 - esc(u, v))(1 - esc(v, w)) \geq 1 - esc(u, v) - esc(v, w)\).

Given that the Markov chain is absorbing with \(A\) the set of absorbing states, the following relations for states \(s \neq t\) are easy to verify:

\[
esc(s, t) = \frac{P^{(s,t)}(s, A)}{P^s(s, t) + P^{(s,t)}(s, A)} = \frac{P^{(s,t)}(s, A)}{1 - P^{A\setminus\{t\}}(s, s)}
\]

\((1)\)

\[a(s) = P^s(s, t)esc(t, s) + P^{(s,t)}(s, A)
\]

\((2)\)

which imply \(P^s(s, t)\mu(s, t) \leq a(s) \leq \mu(s, t)\) and \(a(t)P^s(s, t) \leq a(s)\) \((3)\).

For all these quantities and following ones, we can drop the subscripts and superscripts if there is no ambiguity.

### 2.5 Evaluations

We had extended the values \(r^i : A \rightarrow \mathbb{R}\) on the absorbing states to functions \(r^i\) on all paths in \(\mathcal{H}\). For any stationary strategies \((x, y)\) and players \(i = 1, 2\) extend the definition of \(r^i\) again to a harmonic function \(r^i_{x,y} : S \rightarrow \mathbb{R}\) with \(r^i_{x,y}(s)\) equal to the expected value of \(r^i\) on \(\mathcal{H}\) as determined by \((x, y)\).

For any harmonic function \(r\) on \(S\), and \(p\) a part of or an alternative to the transition from a state \(s\), define \(v^r(p)\) to be the expected value of \(r\) conditioned on the use of \(p\) and no return to the state \(s\), with \(v^r(p)\) defined to be \(r(s)\) if there is return to \(s\) with certainty. If the Markov chain is absorbing and \(g(p) > 0\) then \(v^r(p)\) would be the new harmonic function value for \(s\) if the transition from \(s\) were replaced by \(p\). For every pair of moves \(a \in A^i_1\) and \(b \in A^i_2\) \(v^i_{x,y}(a, b)\) is defined to be \(v^{r^i_{x,y}}(a, b)\) of the part of the transition defined by the pair \((a, b)\) of moves. Likewise define \(v^i_{x,y}(a)\) and \(v^i_{x,y}(b)\) with respect to the pairs \((a, y \in Y^s)\) and \((x \in X^s, b)\), respectively. If \((x, y)\) is absorbing we have the relation

\[
r^i_{x,y}(s) = \frac{\sum_{a,b} x^a_y^b v^i_{x,y}(a, b) g^s_{x,y}(a, b)}{a_{x,y}(s)} = \sum_{a,b} v_{x,y}(a, b) v^i_{x,y}(a, b).
\]

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For $a \in A^*_1$ we have
\[
v^i_{x,y}(a) := \sum_{b \in A^*_2} y^a_{x,y}(a, b) g_{x,y}(a, b) / g^a_{x,y} = \sum_{b \in A^*_2} v_{x,y}(a, b) g_{x,y}(a, b) / \nu^a_{x,y}
\]
and for $b \in A^*_2$ we have
\[
v^i_{x,y}(b) := \sum_{a \in A^*_1} x^a_{x,y}(a, b) g_{x,y}(a, b) / g^a_{x,y} = \sum_{a \in A^*_1} v_{x,y}(a, b) g_{x,y}(a, b) / \nu^a_{x,y},
\]
with both quantities $v^i(s)$ when the quotient is not well defined.

For any harmonic function $r$ on $S$, and $p$, a part of or an alternative to the transition from a state $s$, define $w^r(p)$ to be the expected value of $r$ on the following stage according to the one-time use of $p$ on that stage. We have $w^r(p) = g(p)v^r(p) + (1 - g(p))r(s)$. For any pair of moves $a \in A^*_1$ and $b \in A^*_2$ at $s \in N$ and $i = 1, 2$ $w^i_{x,y}(a, b)$ is the expected value of $r^i_{x,y}$ on the next stage if the players use the pair $a$ and $b$ on the present stage at $s$. For all $b \in A^*_2$ define $w^i_{x,y}(b) := \sum_{a \in A^*_1} x^a_{x,y}(a, b)$ and for all $a \in A^*_1$ define $w^i_{x,y}(a) := \sum_{b \in A^*_2} y^a_{x,y}(a, b)$.

The following is a central lemma concerning the changes in a harmonic function.

**Lemma 2.1:** Let $S$ be the finite state space of an absorbing time homogeneous Markov chain and $r : S \to \mathbb{R}$ a harmonic function. For every non-absorbing $s \in S$ let $p_s$ be an alternative transition at $s$ such that $g(p_s) > 0$. Define a new time homogeneous Markov chain according to the $p_s$. Let $a_s : S \to [0, 1]$ be the absorbing rates corresponding to the new time homogeneous Markov chain and let $r_s : S \to \mathbb{R}$ be a harmonic function with respect to the new transitions such that $r_s$ agrees with $r$ on the absorbing states. If $|v^r(p_s) - r(s)| \leq \delta_s$ and $a_s(s) \geq \epsilon_s g(p_s)$ for $0 < \epsilon_s \leq 1$ and all non-absorbing $s \in S$ (with $g(p_s) = a(s)$ if $p_s$ was the original transition at $s$) then the new Markov chain is absorbing and $|r_s(s) - r(s)| \leq \sum \delta_t / \epsilon_t$ for all states $s$.

**Proof:** The new Markov chain is absorbing because $a_s(s) > 0$ for all $s \in S$. With a start at any state $s_0$, we can bound the change $|r_s(s_0) - r(s_0)|$ by the sum over all states $t \in S$ of the one stage deviation at $t$ multiplied by the expected number of visits to the state $t$. The deviation from one visit to a state $t$ is bounded by $|w^r(p_t) - r(t)|$, and since $1 / a_s(t)$ is the expected number of visits to the state $t$ we have the total deviation bounded by $\sum_t |w^r(p_t) - r(t)| / a_s(t)$. $|w^r(p_t) - r(t)| \leq g(p_t) |v^r(p_t) - r(t)|$ implies $|w^r(p_t) - r(t)| / a_s(t) \leq |v^r(p_t) - r(t)| / \epsilon_t$. \qed
3 Changes in Taboo Probabilities

In all the lemmata of this section, \(S\) is a finite state space of a time homogeneous Markov chain.

3.1 Reaching a State

For the first three lemmata we look at what happens when a fraction of \(P^t(t, s)\) is removed from the transitions at all \(t\) in a set \(T\).

**Lemma 3.1** Let \(s\) and \(t\) be two distinct states and \(A\) and \(B\) two subsets of states such that \(A, B\) and \(\{s, t\}\) are mutually disjoint. Let \(p\) be a part of the transition at \(t\) such that at least positive \(\gamma < 1\) of the transition \(P^{B\cup\{t\}}(t, A)\) goes through \(p\) (meaning that if the complement of \(p\) were removed and replaced by motion that went back to \(t\) on the next stage with certainty then the new quantity for \(P^{B\cup\{t\}}(t, A)\) would be at least \(\gamma\) times the old quantity). If the existing transition at \(t\) were replaced by \(p\) (followed by normalization) and the new transitions were indexed by \(\ast\) then \(P^{B\cup\{t\}}_{\ast}(t, A) \geq \gamma P^{B\cup\{t\}}(t, A)\) and \(P^{B\cup\{s\}}_{\ast}(s, A) \geq \gamma P^{B\cup\{s\}}(s, A)\).

**Proof:** \(P^{B\cup\{t\}}_{\ast}(t, A) \geq \gamma P^{B\cup\{t\}}(t, A)\) is given. If there was never motion from \(s\) to \(t\) or from \(t\) to \(s\) then the inequality \(P^{B\cup\{s\}}_{\ast}(s, A) \geq \gamma P^{B\cup\{s\}}(s, A)\) would also be straightforward. So let us assume that there is some motion in both directions between \(s\) and \(t\), and let \(A'\) be the set \(A\) unioned with all the other states from which there is no motion to either \(s\) or \(t\).

To estimate \(P^{B\cup\{s\}}_{\ast}(s, A)\) let \(b := P^{B\cup\{s,t\}}_{\ast}(s, A), c := P^{B\cup\{A'\cup\{s\}\}}_{\ast}(s, t), d := P^{B\cup\{A'\cup\{t\}\}}_{\ast}(t, s)\) and \(e = P^{B\cup\{s,t\}}_{\ast}(t, A)\). Let \(d_s\) and \(e_s\) stand for the contributions to \(d\) and \(e\) made by the transitions in \(p\), so that \(d_s \leq d\) and \(e_s \leq e\). By assumption we have \(c_s + d_s \frac{b}{b+c} \geq \gamma(e + d_s \frac{b}{b+c})\). We suppose for the sake of contradiction that \(\gamma P^s(s, A) = \gamma (c + \frac{ce_s}{d_s + e_s}) > b + \frac{ce_s}{d_s + e_s} = P^s_s(s, A)\). Re-write as \((d_s + e_s)(be_s + bd_s + ce_s) > (be_s + bd_s + ce_s)(d + e)\) or \(d_s + e_s > d + e\), a contradiction. \(\square\)

**Lemma 3.2**: Let \(T\) and \(A \cup U\) be mutually disjoint subsets of \(S\). If no more than a frequency of \(\gamma P^{T \cup \{u\}}(u, A)\) is removed from the transitions of all \(u \in U \setminus A\) for some fraction \(0 < \gamma < 1/(2|U|)\) and no more than a frequency of \(\gamma\) in the case of \(u \in U \cap A\), followed by normalization, then for all \(x \in S \setminus A\) the new resulting probabilities \(P_{\ast \cup \{x\}}^T(x, A)\) satisfy
$P^{T \cup \{x\}}_s(x, A) \geq (1 - \gamma|U|)P^{T \cup \{x\}}_s(x, A)$ and for every $a \in A$ and $x \notin A \cup T$

(1 - 3|U|\gamma)P^{T \cup \{u\}}_s(a, x) \leq P^{T \cup \{u\}}_s(a, x).

**Proof:** For $U = \emptyset$ there is nothing to prove. Now assume the result for $U \setminus \{u\}$, and let $P_+$ stand for the probabilities where the changes are made in $U \setminus \{u\}$. Since by induction $P^{T \cup \{u\}}_+(u, A) \geq (1 - \gamma|U| + \gamma)P^{T \cup \{u\}}_+(u, A)$, the frequency removal at $u$ is no more than $1 - \frac{\gamma}{1 - \gamma|U| + \gamma}P^{T \cup \{u\}}_+(u, A)$. By Lemma 3.1 applied to the case of only one change at $u$, we have for all $x P^{T \cup \{u\}}_+(x, A) \geq (1 - \frac{\gamma}{1 - \gamma|U| + \gamma})P^{T \cup \{x\}}_+(x, A) \geq (1 - \frac{\gamma}{1 - \gamma|U| + \gamma})(1 - \gamma|U| + \gamma))P^{T \cup \{x\}}_+(x, A) = (1 - \gamma|U|)P^{T \cup \{x\}}_+(x, A).

The second half, if $u \in A$ then it follows by induction because the only way to increase this probability is through the normalization. Otherwise express $P^{T \cup \{u\}}_+(a, x)$ as $P^{T \cup \{u\}}_+(a, x) + \frac{P^{T \cup \{x\}}_+(u, A)}{1 - P^{T \cup \{x\}}_+(u, A)}$. We notice that $1 - P^{T \cup \{u\}}_+(u, u) \geq P^+_s(u, T \cup A \cup \{x\}) \geq P^{T \cup \{u\}}_+(u, A)$, so that the change $1 - P^{T \cup \{u\}}_+(u, u)$ to $1 - P^{T \cup \{u\}}_+(u, u)$ cannot be a decrease by more than a factor of $\gamma/(1 - \gamma|U| + \gamma) \leq 2\gamma$. The rest follows by $(1 - \gamma)P^{T \cup \{u\}}_+(u, x) \leq P^{T \cup \{u\}}_+(u, x)$, (since the only way to increase this probability is through the normalization). \qed

**Lemma 3.3** Let $T$ be a subset of $S$ and let $s$ be a fixed state such that $s$ is reached with positive probability from every $t \in T$. For every $t \in T$ let $q^t$ be a part of the transition at the state $t$ satisfying $f_{qt}P^t_{qt}(t, s) \leq \gamma P^t(t, s)$ where $P^t_{qt}(t, s)$ is the resulting taboo probability if $q^t$ is a replacement transition at $t$. Consider new transitions resulting from the removal of the part $q^t$ at every $t \in T$, followed by normalization. If $|T|\gamma < 1$ then $s$ is also reached with positive probability from all of $T$ after the changes.

**Proof:** We prove by induction on the size of $T$; by Lemma 3.1 the claim holds for $|T| = 1$. With $v \in T$ also fixed, let us assume that there is some state $u \in T$ such that after the changes from a start at $v$ the state $u$ is not reached at all. Whether or not one reaches $s$ from $v$ with the changes cannot not be influenced by any change made at $u$. Therefore by the induction hypothesis, considering changes made in the smaller set $T \setminus \{u\}$, we have our result.

Now assume that with the changes all member of $T$ are reached from $v$. For every pair $t, u \in T$ let $w_t(u)$ be the probability in the original Markov chain with respect to a start at $t$ that $s$ is reached and that the last visit to a
state in $T$ was at the state $u$. Because starting at $t$ rather than at $u$ cannot be a better way to reach $s$ through the state $u$, we have $w_t(u) \leq w_u(u)$. But then there must be a $u \in T$ such that $w_u(u) \geq \frac{1}{|T|} \sum_t w_t(t) \geq \frac{1}{|T|} \sum_t w_u(t)$. This means that at least $\frac{1}{|T|}$ of the original motion $P_u(u, s)$ went directly to $s$ without passing through any other member of $T$ (and therefore after the changes there is still motion from $u$ to $s$). \hfill \Box

The following lemma concerns transitions in two person stochastic games, but can be generalized to any time homogeneous Markov Chain whose transitions are determined by two independent variables.

Lemma 3.4 Let $R$ be a subset of non-absorbing states, $U$ a subset of $R$, and $(x, y)$ a pair of stationary strategies such that there is some motion between all pairs of states in $R$. Let $s, t \in U$ be special states. Assume for every $u \in U \setminus \{s\}$ that no more than a frequency of $\gamma P_u(u, s)$ is removed from $x^u \in X^u$ and no more than a frequency of $\gamma$ from $x^s$, followed by normalization; let $\pi$ stand for the result. Assume for the state $t \in U$ that $P^u_{x,y}(s, t) \geq \epsilon P^u_{x,y}(s, t)$. Let $y^u_*$ be a part of $y^u$ for any $u \in U$ with $f^u_*$ its frequency. Assume for all $u \in U$ and both $z \in \{s, t\}$ that $f^u_0 P^u_{x,(y|y^u_*)}(u, z) \leq \delta P^u_{x,y}(u, z)$ where $(y|y^u_*)$ is the strategy that is $y^u$ when $v \neq u$ and is $y^u_*$ otherwise. Let $\gamma$ stand for the result when $y^u_*$ is removed from $y^u$ for every $u \in U$, followed by normalization. Given that $(1 - 4\gamma|U|) \epsilon > \delta|U|$ with $(\pi, \gamma)$ there is some motion from all states in $R$ to $s$ and also some motion from $s$ to $t$.

Proof: Since the part of $P^u_{x,y}(u, s)$ that was removed cannot exceed $\delta + \gamma$ of the whole, we have from Lemma 3.3 that $s$ is reached from all states $v$ in $R$.

As with the proof of Lemma 3.3 we can assume by induction that all $u \in U \setminus \{t\}$ are reached from $s$ with $\pi$ and $\gamma$. We account for $P^u_{x,y}(s, t)$ by considering the last state visited on the way from $s$ to $t$. For any choice of $(\tilde{x}, \tilde{y})$ let $p_{\tilde{x},\tilde{y}}(u, t) := P^u_{\tilde{x},\tilde{y}}(u, t)$ be the probability of moving from $u$ to $t$ with no other member of $U$ in between. Let $U' := U \setminus \{s, t\}$. We have

$$P^u_{\tilde{x},\tilde{y}}(s, t) = p_{\tilde{x},\tilde{y}}(s, t) + \sum_{u \in U'} p_{\tilde{x},\tilde{y}}(u, t) P_{\tilde{x},\tilde{y}}^{(t, s)}(s, u) \left(1 - P_{\tilde{x},\tilde{y}}^{(s, t, u)}(u, u)\right),$$

since $P_{\tilde{x},\tilde{y}}^{(t, s)}(s, u) / \left(1 - P_{\tilde{x},\tilde{y}}^{(s, t, u)}(u, u)\right)$ is the expected number of times that $u$ is visited before
reaching \( t \) or returning to \( s \), with \( 1 - P^{(s,t,u)}(u,u) = P^u(u,\{s,t\} \cup A) \geq P^u(u,s) \), where \( A \) is the set from which there is no motion to the set \( R \).

Define for all \( u \in U' \) \( e(u) := \frac{p_{x,y}(s,u)}{1 - p_{x,y}(u,u)} \), with \( e(s) = 1 \), and define \( e_u(u) \) correspondingly with respect to \( x \) and \( y \), with \( e_u(s) = 1 \). By Lemma 3.2 we have \((1 - 4\gamma(|U| - 2))e_u(u) \leq e(u)\) for all \( u \in U' \). We can conclude that

\[
\sum_{u \in U \setminus \{t\}} e(u)p_{x,y}(u,t) \geq (1 - 4(|U| - 1)\gamma)P_{x,y}^s(s,t) \geq e(1 - 4(|U| - 1)\gamma)P_{x,y}^s(s,t) = \epsilon \sum_{u \in U \setminus \{t\}} e(u)p_{x,y}(u,t). \tag{4}
\]

Next define \( p_{x,\pi}(u,t) := P_{x,\pi}^t(u,t) \). By recognizing that \( p_{x,y}(u,t)e(u)/P^s(s,t), \) the probability that the last visit to \( U \) was at \( u \in U \) from a start at \( s \), is less than or equal to the probability that the last visit to \( U \) was \( u \) with a start at \( u \) (both according to \((x,y)\)) , we have from the defining condition on \( \pi \) that \( |p_{x,y}(u,t)e(u) - p_{x,\pi}(u,t)e(u)| \leq \delta P^s(s,t) \). After summing over \( U \setminus \{t\} \) we get

\[
\sum_{u \in U \setminus \{t\}} e(u)p_{x,\pi}(u,t) \geq (1 - \delta|U| + \delta) \sum_{u \in U \setminus \{t\}} e(u)p_{x,y}(u,t) \tag{5}.
\]

To show that \( u \) reaches \( t \) for some \( u \in U \setminus \{t\} \), it suffices to show that \( p_{x,\pi}(u,t) + p_{x,\pi}(u,t) > p_{x,y}(u,t) \) for some \( u \in U \setminus \{t\} \). But assuming that \( p_{x,\pi}(u,t) + p_{x,\pi}(u,t) \leq p_{x,y}(u,t) \) for all \( u \in U \setminus \{t\} \), from the above sums in (4) and (5) we must conclude that \( 1 - \delta|U| + \epsilon(1 - 4\gamma|U|) < 1 \), a contradiction to the initial assumption.

\( \square \)

### 3.2 Continuity and Exiting

Because of the unlimited number of stages, taboo probabilities and harmonic functions of time homogeneous Markov chains are not continuous with respect to absolute changes in transition probabilities. However, there is a continuity for relative changes in these transitions. A result of the same spirit but in a different formal context is contained in Freidlin and Wentzell (1984).

**Lemma 3.5** Assume that the transitions \( p^s \in \Delta(S) \) at a subset \( U \) are changed such that for all \( t \in S \), including \( s = t \), the resulting \( p^s(t) \) differs from \( p^s(t) \) by no more than a factor of positive \( \gamma < 1/(2|U|) \) (necessarily
with \( p_s^*(t) = 0 \) if and only if \( p^*(t) = 0 \). Let \( P_{i}^{T}(s, A) \) stand for the resulting taboo probability. For all choices of \( s, T, \) and \( A \) with \( T \cap A = \emptyset \), \( P_{i}^{T}(s, A) \) differs from \( P_{i}^{T}(s, A) \) by a factor of at most \( 4\gamma|U| \). If the original Markov is absorbing then the resulting Markov chain is absorbing and if \( r : S \to \mathbb{R} \) is a harmonic function with respect to the original Markov chain and \( r_\ast \) is the resulting harmonic function that agrees with \( r \) on all the absorbing states then \( |r(s) - r_\ast(s)| \leq 4\gamma|U|M \) for every \( s \in S \), where \( M \) is a bound on the difference between the function values of \( r \) on these absorbing states.

\textbf{Proof:} Let \( U := \{s_1, \ldots, s_N\} \). Let \( P_{i}^{T}(s, A) \) stand for the taboo probability when the changes are made only at the subset \( \{s_1, s_2, \ldots, s_i\} \), and define \( \text{esc}_i(t, s) \) in the same way.

First we claim that for every fixed choice of \( s, T, A \) with \( s \in T \) that \( P_{i}^{T}(s, A) \) and \( P_{i-1}^{T}(s, A) \) differ at most by a factor of \( 2\gamma \). Since both \( P_{i}^{T}(s_i, A) \) and \( P_{i-1}^{T}(s_i, A) \) are expectations over the next stage of some probabilities, we have our claim for \( P_{i}^{T}(s_i, A) \) and a factor of \( \gamma \) by the defining assumption. If \( s \neq s_i \) then we get our result from the same observation and the formula

\[
P_{i}^{T}(s, A) = P_{i}^{T,s_i}(s, A) + P_{i}^{T}(s, s_i)P_{i}^{T,s_i}(s_i, A)/P_{i}^{s_i}(s_i, T \cup B \cup A \cup \{s\}),
\]

where \( B \) is the set of states such that in either the \( i \)th or \( i+1 \)st Markov chain there is no motion to the state \( s_i \) from the set \( B \).

From formula (1) we have \( 1 - \text{esc}_N(t, s) = P_N^s(t, s)/(P_N^s(s, t) + P_N^{s,t}(s, B)) \) and from above that \( 1 - \text{esc}_N(t, s) \) does not differ from \( 1 - \text{esc}(t, s) \) by more than a factor of \( 2\gamma N \). Notice that \( 1 - a(s) \) can be written as the expected value of \( 1 - \text{esc}(t, s) \) on the next stage, and therefore \( 1 - a(s) \) does not differ from \( 1 - a_N(s) \) by more than a factor of \( 2\gamma N \), where \( a_N \) is the resulting absorption rate. This implies that \( a(s) = 1 \) if and only if \( a_N(s) = 1 \) and in this case we have \( P_N^T(s, A) = P_N^{T,s}(s, A) \), \( P_T(s, A) = P_T^{s}(s, A) \), and our result. Given \( a(s) \neq 1 \) then \( P_T^T(s, A) = P_T^{T,s}(s, A)/(1 - a(s)) \) and \( P_N^S(s, A) = P_N^{T,s}(s, A)/(1 - a_N(s)) \) we also have our result. The claim concerning harmonic functions follows by considering \( A \) to be any subset of absorbing states.

Next we define the concept of exit. (Due to the lack of the semi-algebraic analysis, we will be more restrictive in our definition of an exit than Vieille 2000a or Solan 2000.) For any subset \( P \) of non-absorbing states a system of \textit{exits} from \( P \) is a collection of parts of the transitions at the states in \( P \) such that all motion from \( P \) to \( S \setminus P \) must occur through one of these parts.
Each part in the collection is called an exit. Given that the Markov chain is absorbing any subset of non-absorbing states must have a system of exits.

Assume that there is a partition $P$ of the states such that $\{s\}$ is in $P$ for every absorbing $s$ and for every non-absorbing $s \in P \in P$ $q^s \in \Delta(S)$ is the transition defined conditionally by the union of all the exits from $P$ at the state $s$. Let $A$ be the set of absorbing states. For every $P \in P$ let $s_P \in P$ be a representative for the set $P$. We will create two new time homogeneous Markov processes, one by extending the state space and the other by contracting it. These constructions are also in Vieille (2000c).

First we extend the state space. For every $s \in P \in P$, create two new states $s^a$ and $s^b$. Define $S_* := \{s^a \mid s \in A\} \cup_{s \in S \setminus A} \{s^a, s^b\}$, and the corresponding Markov chain will be indexed by $\ast$. The states $\{s^a \mid s \in A\}$ remain absorbing. At $s^a$ with $s \in P \in P$, the motion goes deterministically to $s^b$. At $s^b$ the transition is labeled $p^s \in \Delta(S_*)$. Let $f_s$ be frequency with which $q^s$ is used. Let $p^a$ be the transition defined by $p^a$ conditioned on the non-use of $q^s$, given of course that $f_s \neq 1$. Define $p^a(t^a) = f_s q^s(t)$ and $p^a(t^b) = (1 - f_s) p^a(t)$ (and otherwise zero if $p^a$ is not defined), with $p^a(a) = p^a(a)$ if $a \in A$.

Given that the Markov chain is absorbing, next we contract the state space. Define $S = \{s_P \mid P \in P\}$. A previously absorbing state remains absorbing. For every non-absorbing state $s_P$ let the transition at $s_P$ be induced by the distribution on the next state $t^a$ following $s^b_P$ in the above Markov chain defined on $S_*$. If $t^a$ is absorbing, then $t$ is that next state. If $t^a$ is not absorbing, the $u = u_P$ is the next state with $t \in P' \in P$. Since the Markov chain on $S_*$ is absorbing, modulo events of zero probability the transitions of $S$ are well defined. In a different context (without taboo probabilities) a similar statement to the next lemma was proven by Vieille (2000c).

**Lemma 3.6:** Assume that the Markov chain is absorbing. Let $r$ be a harmonic function on $S$ and $M > 0$ a uniform bound on all differences in the values of $r$. Let $N$ be the number of the $P$ that are not singletons, and let $0 < \delta < \frac{1}{2N}$ be given. Assume for every $P \in P$ and every distinct pair $s,t \in P$ that the probability of moving from $t$ to $s$ without passing through any exit of $P$ is at least $1 - \delta$. The new processes on $S_*$ and $S$ are absorbing and for any pair of subsets $A$ and $T$ that are unions of members of $P$ with $A \cap T = \emptyset$ we have that $P^{A_1}(s, T_*)$ differs from $P^A(s, T)$ by no more than a factor of 20.
4N\delta, where \( B_s := \{ s^a, s^b \mid s \in B \} \) for all subsets \( B \). With \( r_s \) representing the new harmonic function on \( S \) determined by the expected value of \( r \) on the absorbing states and \( r_t \) the same for \( S_t \) we have \( r_s(s^a_R) = r_t(s_R) \) for all representative states \( s_R \) and \( |r_s(s^a) - r(s)| \leq 4MN\delta \) for all \( s \in S \).

**Proof:** Define two new transitions \((\hat{p}^s | s \in S)\) and \((\overline{p}^s | s \in S)\) on \( S \). \( \hat{p}^s \) is determined by the distribution on the next state \( t^s \) in \( S \) from a start at \( s^a \in S \). \( \overline{p}^s \) is defined likewise, however from a start at \( s^b \in S \). The distribution on the states outside of \( P \) with the \( \overline{p}^s \) is the same as with the original transitions \( p^s \) on \( S \). Because of our assumption concerning the avoiding of exits, Lemma 3.5 applies to the difference between \( \hat{p} \) and \( \overline{p} \). The claim for the taboo probabilities follows directly from Lemma 3.5, as does also the claim for the harmonic functions.

Lemma 3.6 works because it is based upon the rare use of an exit. Much more problematic is analysing the consequences of the certain use of an exit. This is the content of Lemma 3.7.

**Lemma 3.7** Assume the context of Lemma 3.6 and that \( p \) is an exit from \( P \) at \( t \in P \) with \( g(p) > 0 \). We have

1) \( |g(p) - g_t(p)| \leq 4N\delta + \delta \),
2) \( g(p) \) and \( g_t(p) \) differ by a factor of no more than \( 4N\delta + \frac{2\delta}{\nu(p)} \),
3) \( \nu(p) \) and \( \nu_t(p) \) differ by a factor of no more than \( 4N\delta + 2\delta + \frac{4N\delta + \delta}{g(p)} \),
4) \( |\nu(p) - \nu_t(p)| \leq 8N\delta + 4\delta \),
5) \( |v^*(p) - v^t(p)| \leq M \min\{8N\delta + \frac{\delta}{g(p)}, 8N\delta + \frac{2\delta}{\nu(p)}\} \).

**Proof:** 1) We define \( \hat{g} \) to be the probability that there is no return to the set \( P \) after using the exit \( p \) in the original Markov chain. From Lemma 3.6 we see that \( \hat{g} \) is within a factor of \( 4N\delta \) of \( g_t(p) \). From the avoiding of exits we get that \( |\hat{g} - g(p)| \leq \delta \), which suffices.

2) By definition \( g(p) \geq \hat{g} \). First we show that \( \text{esc}(u, t) \leq \delta \hat{g}/((1 - \delta)\nu(p)) \) for all \( u \in P \). Define \( w_u \) be the probability that \( p \) will be used before returning to \( u \) from a a start at \( t \) (with \( w_u \leq \delta \) for all \( u \in P \)). Define \( \nu_u \) to be the probability that the last visit to \( P \) is through the exit \( p \) from a start at \( u \); we have \( \nu_u \leq w_u \hat{g}/(w_u \hat{g} + \text{esc}(u, t)) \), which translates to \( \text{esc}(u, t) \leq w_u \hat{g}/\nu_u \). Finally notice that \( \nu_u \) doesn’t differ from \( \nu(p) \) by a factor of more than \( \delta \).

Next we compare \( g(p) \) with \( \hat{g} \). For every \( u \in P \) let \( \lambda_u \) be the probability that there is a return to \( P \) from the use of \( p \) in the original Markov chain.
and that \( u \) is the first member of \( P \) reached. Notice that \( \sum u \lambda_u = 1 - \hat{g} \). We have \( g(p) = \hat{g} + \sum u \lambda_u \text{esc}(u,t) \). This suffices for \( (1 - 2\delta/\nu(p))g(p) \leq \hat{g} \). Now use Lemma 3.6 for the conclusion.

3) By definition \( \nu_*(p) = f_p g_*(p)/a_*(t^h) \) and \( \nu(p) = f_p g(p)/a(t) \). One way to perceive \( a(t) \) is as the reciprocal of the expected number of visits to \( t \) from a start at \( t \). With this perspective by Lemma 3.6 and the avoiding of exits we get that \( a_*(t^h) \) and \( a(t) \) don’t differ by a factor of more than \( 4\delta + \gamma \). This means that if \( g_*(p) \) and \( g(p) \) don’t differ by a factor of more than \( \gamma \) then \( \nu_*(p) \) and \( \nu(p) \) don’t differ by more than a factor of \( \gamma + 4\delta \). Since \( \nu_*(p) \) is also equal to the probability that the last visit to \( P \) starting at \( s^b_P \) in the Markov chain \( S_* \) went through the exit \( p \) we have that \( \nu_*(p) \) is within a factor of \( \delta \) of \( \nu_!(p) \) and therefore \( \nu_!(p) \) and \( \nu(p) \) don’t differ by a factor of more than \( \gamma + 4\delta + 2\delta \). By the same argument as in Part 1 comparing \( g_!(p) \) with \( g(p) \) we get \( |g_!(p) - g(p)| \leq 4\delta N + 2\delta \) and therefore \( g_!(p) \) and \( g(p) \) cannot differ by a factor of more than \( \frac{4\delta N + 4\delta}{g(p)} \) and our conclusion.

4) The argument of Part 2 can be repeated with the Markov chain defined on \( S_* \) instead of the original on \( S \). The quantity \( g(p) \) would be replaced by \( g_*(p) \) and \( \hat{g} \) would be replaced by \( g_!(p) \). We have \( g_!(p) \geq g_!(p) \) and \( g_*(p) = g_!(p) + (1 - g_!(p))\text{esc}_*(s^b, t^h) \).

If \( g(p) \geq g_!(p) \) we need only \( g_!(p) \geq g_!(p) \) and the conclusion of Part 2 to get \( g(p) \geq (1 - 2\delta/\nu(p) - 4\delta N)g_!(p) \). Combined with the arguments from Part 3 we have our goal. On the other hand, if \( g_!(p) \geq g(p) \) we get our result from repeating Part 2 for \( g_!(p) \) and \( g_!(p) \), the same arguments of Part 3, plus the claim that \( (1 - 4\delta N - \delta)\text{esc}_*(s^b_P, t^h) \leq \text{esc}(s_P, t) \).

\( \text{esc}_*(s^b, t^h) \) is no more than \((w + w^2 + \ldots)h_\ast \) where \( w \) is the probability of reaching an exit of \( P \) from \( s^b_P \) before returning to \( t^h \) and the quantity \( h_\ast \) is the expected value of \( g_! \) conditioned on the use of one of these exits. On the other hand we have that \( \text{esc}(s_P, t) \) is at least \( w\hat{h} \) where \( \hat{h} \) is the probability of no return to the set \( P \) in the original Markov chain conditioned on the use of one of these exits. From Lemma 3.6 we have that \( \hat{h} \) and \( h_\ast \) differ by no more than a factor of \( 4\delta N \). That \( w \leq \delta \) completes the proof of the claim.

5) From the proof of Part 1 we had that \( \hat{g} \geq (1 - \delta/\nu(p))g(p) \) and from Part 2 that \( \hat{g} \geq (1 - 2\delta/\nu(p))g(p) \). The rest follows from Lemma 3.6.

Part 4 of Lemma 3.7 is remarkable because the sum of \( \nu \) over all transitions in a set \( P \) will be |\( P \)| rather than something close to one.
3.3 Polarization

The process described below, of changing the transitions through a convex combination of two transitions, one giving a higher value and the other giving a lower value of a harmonic function, with the convex combination yielding the same value, we call polarization.

Lemma 3.8 Let $s$ and $t$ be two non-absorbing states of an absorbing Markov chain.

(i) Let $p$ be a part of the transition at $t$ such that $\nu(p) \geq \epsilon > 0$.

(ii) Let $p$ be a replacement transition at $t$ such that $g(p) \geq \epsilon$.

(iii) Let $p$ be a transition at $t$ that is a convex combination of transitions as described in (i) and (ii).

In all three above cases, if we replace the transitions at $t$ by $p$, in the case of (i) or (iii) using normalization, the resulting process is absorbing and the absorption rate of $s$ is at least $\epsilon$ times what is was before the changes were made.

Proof: Let $b$, $c$, $d$ and $e$ stand for the same quantities as in the proof of Lemma 3.1, with $A$ the set of absorbing sets and $B$ the empty set.

(i) It follows from Lemma 3.1.

(ii) Let $a_*, d_*$ and $e_*$ be the corresponding quantities when $p$ is the transition at $t$. We assume that $e_* + d_* \frac{b}{b+c} \geq \epsilon$. Suppose for the sake of contradiction that $\epsilon(b + \frac{ce}{d+e}) = \epsilon a(s) > a_*(s) = b + \frac{ce}{d+e}$. Then we have $be_* + ce_* + bd_* \geq (b+c)\epsilon \geq \epsilon(b + \frac{ce}{d+e}) > \frac{bd_* + be_* + ce_*}{d_* + e_*}$. This implies $d_* + e_* > 1$, also a contradiction.

(iii) First we must assume that $b < \epsilon a(s)$, since otherwise there would be nothing to prove. Let $a_i$, $d_i$ and $e_i$ for $i = 1, 2$ stand for the resulting probabilities from (i) and (ii), respectively, and after normalization in the case of (i). With the convex combinations $\tilde{d} := \lambda d_1 + (1 - \lambda) d_2$ and $\tilde{e} := \lambda e_1 + (1 - \lambda) e_2$ being the new transition quantities, we have that our desired result is equivalent to $\frac{\tilde{e}}{\tilde{e} + \tilde{d}} \geq \epsilon \frac{\epsilon}{\epsilon + \tilde{d}} + \frac{eb - \tilde{e}}{\epsilon}$. But this follows from (i), (ii), and the fact that $\frac{\tilde{a}}{y_1} \geq z$ and $\frac{\tilde{a}}{y_2} \geq z$ implies $\frac{\lambda x_1 + (1 - \lambda) x_2}{\lambda y_1 + (1 - \lambda) y_2} \geq z$ for all non-negative quantities $x_i, y_i, z$ and $0 \leq \lambda \leq 1$. □

Proposition 3.9 Let $r^1$ and $r^2$ be two harmonic functions, and we assume that the Markov chain is absorbing. Let $N$ be the number of non-absorbing states. Let $1$ be a uniform bound on all differences in the values of $r^1$ and
r^2. Let w^1, w^2, v^1, and v^2 stand for w^{r^1}, w^{r^2}, v^{r^1}, and v^{r^2}, respectively. Let 1/2 > \epsilon > \delta > \gamma > 0, with \delta < \frac{\epsilon^3N}{2N(3N)^2}. Let p^*_s be a part of the transition at s such that w^2(p^*_s) \leq r^2(s) - \epsilon (including the possibility that p^*_s is empty). Assume that if \nu(p^*_s) \geq \gamma then there is an alternative transition p_s at s such that w^2(p_s) \leq r^2(s) - \epsilon, |v^1(p_s) - r^1(s)| \leq \delta, and there exists another part q_s of the transition at s such that q^d_s, the complement of the union of q_s with p^*_s, satisfies (v^2(q^d_s) - r^2(s)) \nu(q^d_s) \leq N\delta/\epsilon. For every subset T \subseteq \{s | \nu(p^*_s) \geq \gamma, w^2(q_s) > r^2(s)\} define a new time homogeneous Markov chain by the transitions at t \in T defined by \lambda p_t + (1 - \lambda)q_t with \lambda satisfying \lambda w^2(q_t) + (1 - \lambda)w^2(q_t) = r^2(t) and furthermore for every \nu \in S\setminus T the part p^*_s is discarded, followed by normalization. Let the subscript T stand for the quantities determined by the new transitions with the changes in T.

**Conclusion:** There is a subset T \subseteq \{s | \nu(p^*_s) \geq \gamma, w^2(q_s) > r^2(s)\} such that the new process is absorbing and for both i = 1, 2 and all s \in S \ | _i r^2(s) - r^i(s) | \leq \epsilon

**Proof:** First we consider what happens when the changes are made only at a set T (meaning that the part p^*_s is kept in for \nu \notin T), which we will label with T, *. Because r^2 remains a harmonic function after the changes are made and there is always a positive probability at all states in T that the harmonic function drops by \epsilon, the resulting time homogeneous Markov chain is absorbing with r^2_{T,*}(s) = r^2(s) for every s \in S.

Next we must determine which subset T will be chosen. Choose any t_1 such that \nu(p^*_{t_1}) \geq e^2/2N, and put t_1 in T. If there exists no such t \in S then let T be the empty set. At any set T with |T| = k - 1 formed so far, put into T any t_k such that \nu_{T,*}(p^*_{t_k}) \geq e^2/2N, and stop if there is no such new state t_k.

**Claim:** For any set T that has been already chosen and any t \notin T that could be added to T we have \nu_{T,}(t,u) \geq \frac{e^2}{3N}a_{T,*}(u) \geq \frac{e^3}{3N}a_{T,*}(u) for all u \in S, g_{T,*}(q_t) \geq \frac{e^3}{3N}g_{T,*}(q_t) and w^2(q_t) > r^2(t).

**Proof of Claim:** Assume that t will be added to T. Look at the transition \nu_{T,*}(q^d_t) and the identities w^2_{T,*}(q^d_t) - r^2_{T,*}(t) = w^2(q^d_t) - r^2(t) = (\nu_{T,*}(q^d_t) - r^2_{T,*}(t))g_{T,*}(q^d_t) from the fact that r^2 remains the harmonic function. Consider the definitions \nu_{T,*}(q^d_t) = f_{q^d_t}g_{T,*}(q^d_t)/a_{T,*}(t) and \nu(q^d_t) = f_{q^d_t}g(q^d_t)/a(t); they show that the new absorption rate determines alone the new value (\nu_{T,*}(q^d_t) - r^2_{T,*}(t))\nu_{T,*}(q^d_t). From the induction assum-
tion we must conclude that $\nu_{T,\ast}(q^T_t)(v^2_{T,\ast}(q^T_t) - r^2(t)) \leq \frac{(3N)|T|}{\epsilon^3} \nu(q^T_t)(v^2(q^T_t) - r^2(t)) \leq \frac{(3N)|T|}{\epsilon^3} \frac{N \delta}{\epsilon} < \epsilon^3/6N$. If $q^T_t$ is the union of $q^T_t$ with $p^T_t$ from $\nu_{T,\ast}(p^T_t) \geq \epsilon^2/2N$ and $w^2_{T,\ast}(p^T_t) \leq r^2(t) - \epsilon$ we get that $\nu_{T,\ast}(q^T_t)(v^2_{T,\ast}(q^T_t) - r^2(t)) \leq -\frac{\epsilon^3}{3N}$, which implies that $w^2(q_t) > r^2(t)$ and $\nu_{T,\ast}(q_t) \geq \epsilon^3/(3N)$.

Next suppose for the sake of contradiction that $g_{T,\ast}(q_t) < \frac{\epsilon^3}{3N}(3N)^{-1}|T|^{-1}g(q_t)$.

Since $\nu(q_t) = f_{q_t}g(q_t)/a(t)$, $\nu_{T,\ast}(q_t) = f_{q_t}g_{T,\ast}(q_t)/a_{T,\ast}(t)$ and $\nu_{T,\ast}(q_t) \geq \epsilon^3/(3N)$, by the induction assumption we would be forced to accept $\nu(q_t) > 1$, an impossibility.

By Lemma 3.8 we have our claim on the absorbing rates for all states other than $t$. For the state $t$ we have $g_{T,\ast}(q_t) \geq f_{q_t}g_{T,\ast}(q_t)/a_{T,\ast}(t)$ and $\nu_{T,\ast}(q_t) \geq \epsilon^3/(3N)$. With $g_{T,\ast}(p_t) \geq \epsilon$ our claim is proven.

With the claim we conclude from Lemma 2.1 that $|r^1_{T,\ast}(s) - r^1(s)| \leq \frac{(3N)|T|}{\epsilon^3} \delta N \leq \epsilon/2$ for all $s \in S$.

Next, we must show that it is impossible for any state $s$ not polarized to satisfy $\nu_{T,\ast}(p^*_{s}) \geq \epsilon^2/N$. This holds for all states with $\nu(p^*_{s}) \geq \gamma$, by construction. Let’s assume that $\nu(p^*_{s}) < \gamma$; this means that the probability of ever using $p^*_{s}$ in the original Markov chain cannot exceed $\gamma/\epsilon$. But by the above claim we know additionally that the probability of using $p^*_{s}$ in the altered Markov chain indexed by $T,\ast$ cannot exceed $\frac{\gamma}{\epsilon} \frac{(3N-1)^{N-1}}{(3N)^{N-2}} < \epsilon^2/2N$.

Next we must consider the influence of the removed $p^T_t$ in the above Markov chain indexed by $T,\ast$. For any $s$ with $\nu_{T,\ast}(p^*_{s}) \leq \epsilon^2/2N$ the chance of ever using the transition $p^T_t$ cannot exceed $\epsilon/2N$, and so they cannot contribute an average of more than $\epsilon/2$ to either the function $r^1$ or $r^2$. \qed

4 From Markov Chains to Equilibria

4.1 Application of the Doob-Kolmogorov Inequality

We must prove Proposition 4.2, a cornerstone of our analysis.

**Lemma 4.1:** Let $X$ be the finite state space of a time homogeneous Markov chain with probability transitions $(p^x \in \Delta(X) \mid x \in X)$. Let $v : X \to \mathbb{R}$ be a harmonic function and let $M > 0$ be a bound for the maximal difference between all values of $v$.

For every $x \in X$ define the non-negative quantities $w(x)$ by $w(x) =$
By taking the limit as \( Q \) goes to infinity and \( \delta \leq |r(y)| \) we get

\[
\text{Probability} \left( \max_{i<Q} |R_i| > \epsilon \right) < \frac{1}{\epsilon^2} \mathbf{E} \left( \sum_{i<Q, y \in Y_{x_{i-1}}} p^{x_{i-1}}(y)r(y)^2 \right).
\]

By taking the limit as \( Q \) goes to infinity and \( \delta \leq |r(y)| \) we get

\[
\text{Probability} \left( \max_{i<\infty} |R_i| > \epsilon \right) < \frac{1}{\epsilon^2} \mathbf{E} \left( \sum_{i<\infty, y \in Y_{x_{i-1}}} p^{x_{i-1}}(y)r(y)^2 \right) \leq \delta \frac{1}{\epsilon^2} \mathbf{E} \left( \sum_{i<\infty, y \in Y_{x_{i-1}}} p^{x_{i-1}}(y) |r(y)| \right) = \delta \frac{1}{\epsilon^2} \mathbf{E} \left( \sum_{i<\infty, y \in Y_{x_{i-1}}} \tilde{w}(x_{i-1}) \right).
\]

Since by the triangle inequality \( \tilde{w}(x) \leq w(x) \) for all \( x \), we have

\[
\text{Probability} \left( \max_{i<\infty} |R_i| > \epsilon \right) < \delta \frac{1}{\epsilon^2} \mathbf{E} \left( \sum_{i<\infty, y \in Y_{x_{i-1}}} w(x_{i-1}) \right),
\]

\[\sum_{y \in X} p^x(y)|v(y) - v(x)|. \] Let \( n \) be the number of states \( x \) such that \( w(x) > 0 \). For any path \( p = (x_0, x_1, x_2, \ldots) \) in \( X \) define \( w(p) = \sum_{i=0}^{\infty} w(x_i) \).

**Conclusion:** The expected value of the function \( w \) does not exceed \( Mn \).

**Proof:** We isolate the problem, handling each state \( x \) separately. Since \( |v(y) - v(x)| \) is always less than or equal to \( M \) times \( \text{esc}(y,x) \), we have that \( w(x) \leq a(x)M \). Therefore the part of the sum that comes from visits to \( x \) does not exceed \( a(x)M \sum_{i=0}^{\infty} (1 - a(x))^i = M \).

**Proof of Proposition 4.2** (as stated in the introduction):

Define the random variable \( r_i \) on the odd steps \( i \) to be \( v(y_i) - v(x_{i-1}) \), and \( R_i \) to be the sum of the \( r_k \) for odd \( k \leq i \). For \( y \in Y_x \) define \( r(y) \) to be \( v(y) - v(x) \).

Define a new quantity, \( \tilde{w}(x) := \sum_{y \in Y_x} p^x(y)|v(y) - v(x)| \). Let \( w(x) \) be the old quantity on the Markov chain from Lemma 4.1 defined only on the \( X \), – we ignore the visits to the \( Y_x \) sets, and consider only the motions from \( X \) to \( X \).

The Doob submartingale inequality states that if \((S_i \mid i = 0, 1, \ldots, n)\) is a martingale with zero expectation then for every \( n \geq 0 \), positive value \( c > 0 \) and exponent \( p \geq 1 \) the probability that \( \max_{i \leq n} |S_i| > c \) is less than \( \mathbf{E}(|S_n|^p)/c^p \) (Williams 1991, Section 14.6). Since the martingale property implies that \( \mathbf{E}(S_n^2) \) is equal to the sum over all the stages \( 1 \leq i \leq n \) of \( E(s_i^2) \), where \( s_i \) is the change in value between the \( i - 1 \)st stage and the \( i \)th stage, we have for every finite even and positive \( Q \)

\[
\text{Probability} \left( \max_{i<Q} |R_i| > \epsilon \right) < \frac{1}{\epsilon^2} \mathbf{E} \left( \sum_{i<Q, y \in Y_{x_{i-1}}} p^{x_{i-1}}(y)r(y)^2 \right).
\]
and by Lemma 4.1 this is no more than $\delta M n / \epsilon^2$. So with $\epsilon \leq 1/2$, we have our result from the size of $\delta$.

The problem of extending Proposition 4.2 to Markov chains that are not time homogeneous (or have countably many states) lies with Lemma 4.1 and not in the proof of Proposition 4.2.

The following corollary relates the above work on Markov chains to our two-person stochastic games. Because the application of this corollary involves an altered state space, this result should be understood in an abstract way.

**Corollary 4.3:** Let $(x, y) \in X \times Y$ be stationary absorbing strategies. Assume that
1) for both players $k = 1, 2$ and $s \in S$ $r_{x,y}^k(s)$ is greater than $j_z^k(s) - \epsilon$ with $z = x$ if $k = 2$ and $z = y$ if $k = 1$, and that
2) for both player $k = 1, 2$ and all moves $c$ used with positive probability with $(x, y)$ by Player $k$ the value $w_{x,y}^k(c)$ is within $\delta$ of $r_{x,y}^k(s)$.

**Conclusion:** For any positive $\epsilon < 1/2$ if $\delta$ is no more than $\frac{\epsilon^3}{n}$ then the strategies $(x, y)$ generate a $4\epsilon$-equilibrium of the stochastic game.

**Proof:** We define the following strategy for Player $k$. For every starting point $s_0 \in S$ let $n_{s_0}$ be large enough such that with a start at $s_0$ and the play according to $(x, y)$ the probability that there is no absorption before the $n_{s_0}$th stage is less than $\epsilon/10$. Let $s_0, s_1, \ldots$ be any sequence of states reached in the game and for both $k$ let $c_0^k, c_1^k, \ldots$ be the sequence of moves made by Player $k$. For $k' \neq k$ as long as $\sum_{i=0}^l (w_{x,y}^{k'}(c_i) - r_{x,y}^{k'}(s_i)) \leq \epsilon$ and the stage $l$ does not exceed $n_{s_0}$ and Player $k'$ never chooses $c_i^{k'}$ outside of the support set of his stationary strategy, then Player $k$ continues to act according to his stationary strategy. As soon as one of the above conditions is violated at some stage $l$ then on the next stage $l + 1$ both players punish each other according to the functions $c_1 + \epsilon$ and $c_2 + \epsilon$. (The mutual punishment is necessary because otherwise a player could intentionally prolong the game with an interest in punishing the other player. The result can be extended to multi-player stochastic games if it can be determined who should punish whom in all situations!) That no player $k$ can obtain an expected payoff more than $2\epsilon$ above the function $r^k$ by choosing a different strategy is self explanatory. That punishment occurs before absorption with probability no more than $2\epsilon$ if both players adhere to the suggested strategies follows from
Proposition 4.2.

4.2 Situations

Next we create an expanded state space from the original state space through partitions of the histories. For every $s \in S$ let $P^s$ be a partition of $H^s$. Define $\hat{S}$ to be the disjoint union $\bigcup_{s \in S} P^s$. For every $t \in \hat{S}$ let $b(t) \in S$ be the member of $S$ such that $t \in P^{b(t)}$. A member of $\hat{S}$ we call a situation. We define the situations $\hat{S}$ to be normal if and only if the next $u \in \hat{S}$ following a $t \in \hat{S}$ is determined by the situation $t$, the choice of moves by the players at $t$, and the next $s \in S$ with $b(u) = s$. Normalcy implies that one can define a stochastic game on the situations as a new state space.

Corollary 4.4: Let the situations $\hat{S}$ be normal, let absorbing stationary strategies $(x, y) \in \prod_{s \in \hat{S}} \Delta(A_1^{b(s)}) \times \prod_{s \in \hat{S}} \Delta(A_2^{b(s)})$ be defined on the situations $\hat{S}$, with $\hat{r}_k : \hat{S} \to \mathbb{R}$ the expected payoff for Player $k$ as determined by the above stationary strategies and the functions $r^k$ on the absorbing states and $\hat{w}_k$ the corresponding expected value of $\hat{r}_k$ on the next stage. Assume that

1) for every $s \in \hat{S}$ $\hat{r}_k(s) \geq j^k_k - \epsilon$ where $z = x$ if $k = 2$ and $z = y$ if $k = 1$

2) for every move $c$ used with positive probability at a situation $s$ by Player $k$ $|\hat{w}_k^c(c) - \hat{r}_k^c(s)| \leq \delta$.

If $\delta$ is no more than $\frac{\epsilon}{|\hat{S}|}$ then these stationary strategies generate a $4\epsilon$-equilibrium of the original stochastic game.

Proof: Because a stochastic game is defined by the normality of $\hat{S}$ and the conditions of Corollary 4.3 are preserved, the result follows by Corollary 4.3.

4.3 First Main Theorem

For any subset $R \subseteq \mathcal{N}$ and a state $s \in R$, a pair $a \in A_1^s$ and $b \in A_2^s$ of moves is called a primitive exit from the set $R$ if with positive probability there is motion from $s$ to $S \setminus R$ using the pair $a$ and $b$. By the definition of $\rho$, any use of a primitive exit at $s$ results in a probability of at least $\rho$ of reaching the complement of $R$.

For every subset $B$ of Player Two moves in a set $R$ we define a $B$ exit (or
simply exit if there is no ambiguity) from $R$ to be any pair $(a, b)$ of moves at an $s \in R$ such that $(a, b)$ is already a primitive exit from $R$ or $b \in B$. Let $E^B(R)$ stand for the set of all $B$ exits from $R$.

Define $B^2_{x,y}(s)$ to be those moves of Player Two at the state $s$ with $w^2_{x,y}(b) \leq r_{x,y}(s) - \gamma$, and let $B^\gamma_{x,y}(R)$ be the union of all the $B^2_{x,y}(s)$ for all $s \in R$. For every $s \in \mathcal{N}$ define $z^2_{x,y}(s)$ to be $\sum_{b \in B^2_{x,y}(s)} v^b$. For any subset $R \subseteq \mathcal{N}$ define $z^\gamma_{x,y}(R) := \sum_{s \in R} z^2_{x,y}(s)$.

For any stationary strategy $x \in X$ (or $y \in Y$) define a simplication of $x$ to be another stationary strategy $\tilde{x} \in X$ obtained from $x$ by dropping the use of certain moves, followed by normalizing what remains. Call the simplification a $\gamma$-simplication if the frequency of the moves removed did not exceed $\gamma$. The simplication is within a set $T$ of states if changes were made only within the set $T$.

**Theorem 1:** Assume for every choice of positive $1/2 > \epsilon > \bar{\epsilon} > \tilde{\epsilon} > \bar{\epsilon} > 0$ with $\bar{\epsilon} < \epsilon^3/(50|\mathcal{N}|)$, $\tilde{\epsilon} < \frac{\bar{\epsilon}}{5(3|\mathcal{N}|)^2|\mathcal{N}|}$ and $\bar{\epsilon} < \bar{\epsilon}/40|\mathcal{N}|$ that

1) there are absorbing stationary strategies $(x, y) \in X \times Y$ with
   a) $r^2_{x,y}(s) \geq j^2_x(s) - \epsilon/2$ for all $s \in \mathcal{N}$,
   b) $r^1_{x,y}(s) \geq j^1_y(s) - \epsilon/2$ for all $s \in \mathcal{N}$, and
   c) for every move $a$ of Player One used in $x$ with positive probability at $s$ we have $|w^1_{x,y}(a) - r^1_{x,y}(s)| \leq \bar{\epsilon}$,

2) a partition $\mathcal{R}$ of a subset $P \subseteq \mathcal{N}$ and for every $R \in \mathcal{R}$ a set $B_R$ of Player Two moves in $R$ containing $B^\gamma_{x,y}(R)$ such that
   a) $\forall s \not\in P \ s.t. \ z^{2,5\gamma}_{x,y}(s) < \bar{\epsilon}$ and
   b) for every distinct $s, t \in R \in \mathcal{R}$ the probability of reaching $s$ from $t$ before using a member of $E^{B_R}(R)$ is at least $1 - \gamma^*$ with $\gamma^* := \bar{\epsilon}/(40|\mathcal{N}|)$,

and for any $R \in \mathcal{R}$ if $z^{2,5\gamma}_{x,y}(R) \geq \bar{\epsilon}$ then there is a special subset $D_R \subseteq R$, a representative $s_R \in D_R$ and

3) an $\bar{\epsilon}$ simplication $y_R$ of $y$ within $R$ created by removing the set $B_R$ of moves such that
   a) $w^2_{x,y}(b) \leq r^2(s)$ for every $b \in B_R \cap A^2_x$ and
   b) $|r^1_{x,y}(s_R) - r^1_{x,y}(s_R)| \leq \tilde{\epsilon}$,

4) $\bar{\epsilon}$-simplications $(x_C, y_C)$ of $(x, y)$ within $D_R$ such that with $(x_C, y_C)$ the play never leaves the set $D_R$ and from any state in $D_R$ all other states in $D_R$ are reached with probability one, and

5) a strategy $y_D$ created from $y_C$ by adding to $y_C$ in the set $D_R$ small probabil-
ities of using a subset of Player Two moves \( V_R \) used in \( D_R \) with \( V_R \subseteq B^2_{x,y}(R) \) and a real positive value \( \xi_R \leq r_{x,y}^2(s_R) - 2.4\pi \) such that

a) with \((x_C, y_D)\) for every pair \( s, t \in D_R \) the probability of reaching \( s \) from \( t \) before using a member of \( V_R \) is at least \( 1 - \gamma^* \)

b) \( \xi_R \geq j^2_x(t) - \epsilon \) for all \( t \in D_R \),

c) for all moves \( b \in V_R \) \( |w_{x,y}^2(b) - \xi_R| \leq \tilde{\epsilon} \), and

d) \( |r_{x_C, y_D}^1(s) - r_{x,y}^1(s)| \leq \tilde{\epsilon} \).

**Conclusion:** With the assumption that Player One can send transition independent signals, the stochastic game has approximate equilibria.

**Proof:** We define the set \( B \) of Player Two moves to be \( \bigcup_{R \in \mathcal{R}} B_R \cup_{s \in P} B_s^s \), and define the exits to be the \( B \) exits. Let the corresponding state spaces \( S_i \) and \( S_t \) from Lemma 3.6 be induced by \((x, y)\) and the partition \( \mathcal{R} \cup \{\{s\} \mid s \notin P\} \). For every \( s_R \in S_t \) let \( p_{s_R}^s \) be the transition at \( s_R \) in \( S_t \) induced by the Player Two moves in \( B^2_{x,y}(R) \). For every \( R \in \mathcal{R} \) define \( p_R \) to be the alternative transition from \( s_R \) in \( S_t \) induced by the Player Two moves \( V_R \) according to \((x_C, y_D)\). Define \( q_R^d \) to be the transition induced by the moves in \( B_R \), and define \( q_R^d \) so that \( q_R^d \) is the disjoint union of \( q_R^d \) with \( p_R \).

We will confirm the conditions of Proposition 3.9 on the state space \( S_t \), with \( 2.4\pi, 2\tilde{\epsilon}, \) and \( 2\tilde{\epsilon} \) the quantities \( \epsilon, \delta \), and \( \gamma \) of that lemma, respectively.

First, by Lemma 3.6 the Markov chain on \( S_t \) is absorbing. For \( i = 1, 2 \) let \( r_i^1 : S_t \rightarrow \mathbb{R} \) be the harmonic function that agrees with the function \( r_i^1 \) on the absorbing states. If \( \nu_i(p_{s_R}^s) \geq 2\tilde{\epsilon} \) then by Lemma 3.7 \( z^2_{x,y}(R) \geq 3\tilde{\epsilon}/2 \) and if \( s \notin P \) then \( \nu_i(p_i^s) \leq 1.1z^2\pi_{x,y}(s) \leq 1.1\tilde{\epsilon} \). By Lemma 3.6 we have for every representative \( s_R \) that \( r_i^1(s_R) \) is within \( 4\gamma^*|\mathcal{N}| \) of \( r_i^1(s_R) \). Equally important, Lemma 3.7 implies that \( w_{x,y}^2(p_{s_R}^s) \leq r_i^1(s_R) - 2.4\pi \), and \( |w_i^1(p_R) - r_i^1(s_R)| \leq 11\tilde{\epsilon}/10 \). Since \( q_R^d \) is induced by some \( B_R \) moves by Lemma 3.7 and Condition 3a we have \( (w_i^1(q_R^d) - r_i^1(s_R))\nu_i(q_R^d) < 2\tilde{\epsilon} \).

Left to confirm is that \( |v_i^1(q_R^d) - r_i^1(s_R)| \leq 2\tilde{\epsilon} \). We apply Lemma 3.6 to the pair \((x, y_R)\) and the transitions it induces on \( S_t \). Since the avoiding of exits by \((x, y)\) implies the same for the pair \((x, y_R)\), we have that \( \mathcal{L}^1(s_R) - r_{x,y\mathcal{R}}^1(s_R) \leq 4\gamma^*|\mathcal{N}| \), where \( \mathcal{L}^1 \) is the harmonic function induced by \((x, y_R)\) on \( S_t \). \( \mathcal{L}^1(s_R) \) is equal to \( v_i^1(q_R) \). With the given \( |r_{x,y\mathcal{R}}^1(s_R) - r_{x,y}^1(s) \leq \tilde{\epsilon} \) and the above relation of \( r_{x,y}^1 \) to \( r_{x,y}^1 \) we are done establishing the conditions of Proposition 3.9.

We apply Proposition 3.9 to \( S_t \) with \( T \) the subset of \( \mathcal{R} \) that has been
polarized. We conclude that the new harmonic functions \( \tilde{r}^i_T := (r^i_T) T \) on \( S^i \) satisfy \(|\tilde{r}^i_T(s) - r^i_{x,y}(s)| \leq 3\epsilon \) for all \( s \not\in P \) and \(|\tilde{r}^i_T(s_R) - r^i_{x,y}(s_R)| \leq 3\epsilon \) for all \( R \in \mathcal{R} \).

Next we define the situations \( \hat{S} \), with one, two, or three situations defined for each original state in \( S \). For any \( s \not\in P \) or for \( s \in R \in \mathcal{R} \) with \( R \not\in \mathcal{T} \) not polarized there is only the situation \( s^e \) (including the case of absorbing states). We always start the game at an \( s^e \). At any situation \( s^e \) for \( s \not\in P \) or \( s \in R \not\in \mathcal{T} \) the players perform \((x^*, \tilde{y}^*)\) where \( \tilde{y}^* \) is the \( \gamma^* \) simplication of \( y^* \) resulting from the removal of all Player Two moves in \( B_{x,y}(s) \). If \( s \) is in a polarized \( R \in \mathcal{T} \) and is not the representative \( s_R \) the players perform \((x, y_R)\). Following any \( s^e \) other than \( s = s_R \) the next situation is a \( t^e \), where \( t \) is the next state in \( S \). Also following the performance of an exit, no matter what the situation was on the previous stage, if \( t \in S \) occurs on the following stage then the next situation is also \( t^e \). This means that only motion inside of an \( R \in \mathcal{T} \) involves situations other than those with the subscript \( e \).

At any \( s \in R \in \mathcal{T} \) there is either two situations \( s^e \) and \( s^f \) if \( s \not\in D_R \) or three situations \( s^e, s^f, \) and \( s^g \) if \( s \in D_R \). For such an \( R \in \mathcal{T} \) let \( \lambda_R \) be the quantity determined by the application of Proposition 3.9 to the transitions on \( S^i \). Since Player One can send signals, for every \( s_R \in D_R \) for a polarized \( R \in \mathcal{T} \) we associate one of every pair of her moves with the symbol \( f \) and the other with the symbol \( g \). If \( s_R^e \) is the present situation then with probability \( \lambda_R \) Player One chooses a move associated with the symbol \( g \) and with \( 1 - \lambda_R \) a move associated with the symbol \( f \); in both cases the players perform \((x_C, y_C)\). (Because all moves are paired, we can modify \( x_C \) to use only those moves corresonding to \( f \) or only moves corresonding to \( g \) without changing the transition probabilities in the space \( S^i \).) If \( t \) is the next state and a move corresponding to \( f \) was used, then \( t^f \) is the next situation; otherwise the next situation is \( t^g \). At any \( s^f \) with \( s \in R \in \mathcal{T} \) the play continues according to \((x, y_R)\), always to a next situation \( t^f \) if there was no use of an exit. On the other hand, from any \( s^g \) with \( s \in D_R \) the motion follows \((x_C, y_D)\), and unless a move from \( V_R \) is used the next situation is a \( t^g \), necessarily with \( t \in D_R \).

Define \( \hat{r}^i \) to be the harmonic function on \( \hat{S} \) determined by the above defined stationary behavior and \( \hat{r}^i = r^i \) on the absorbing states. Given the above conditions, to apply Corollary 4.4 it suffices that neither player \( i \) can change the expected value of \( \hat{r}^i \) by more than \( 10\epsilon \) at any one stage. With
the role of the $\xi_R$ we need only show that $\tilde{r}^i$ is within $\tau$ of $r^i$ on all the $s_R$ and the $s \not\in P$. To do this, we introduce two new transitions defined on $S$, indexed by $\circ$ and $\wr$. $p_i$ and $p_s$ are identical on states $s$ that are not in a polarized $R$, and then it is that induced by the behavior at the situation $s^e$. At $s$ in a polarized $R \in T$ $p_i^s$ is the distribution determined by the next situation $t^e$ following the situation $s^e$. $p_s^s$ is determined by the next situation $t^e$ conditioned on having reached either $s^e_R$ or $s^e_{\not R}$ before any exit was performed. The $p_s^s$ transitions generate harmonic functions $r^i_s$ that are identical to $r^i$ on the $S^\sharp$, and the $p_i$ transitions generate harmonic functions $r^i_{\circ}$ that are identical to $\tilde{r}^i$ on the subset $\{s^e | s \in S\}$. Because $\lambda_R$ cannot be greater than $1 - 2\tau$ and the probability from a situation $s^e$ that an exit from the stationary strategies $(x, y_R)$ is used before getting to $s^e_R$ is no more than than $\gamma^*$, for every $s, t \in S$ the transition probability $p_i^s(t)$ does not differ by more than a factor of $\gamma^*/\tau$ from $p_s^s(t)$. Finally Lemma 3.5 implies that the functions $r^i_s$ and $r^i_{\circ}$ do not differ by more than $4\gamma^*|N^\circ|/\tau < \tilde{\epsilon}$. \hfill \Box

5 The auxiliary game

The main issue is to define the “correct” discounted evaluation of Player Two, since, as shown in Solan (2000), a naive definition of his discounted evaluation does not prove equilibrium existence when there are a multitude of non-absorbing states. We assume that positive $\epsilon$ and $\tau$ have been fixed.

5.1 The function $\xi$

Let $b$ be any move of Player Two at a state $s \in N$. For any $(x, y) \in X \times Y$ define

$$
g^b_{x,y} = \begin{cases} 
1 & g^b_{x,y} \geq \tau \\
g^b_{x,y}/\tau & g^b_{x,y} < \tau.
\end{cases}
$$

Define the auxiliary absorption rate by $\tilde{a}_{x,y}(s) = \sum_{b \in B} y^x_b g^b_{x,y}$. Note that $a(s) \leq \tilde{a}(s) \leq a(s)/\tau$.

Define $\tilde{v}^2(b) = (1 - g^b_{x,y}) r^2(s) + g^b_{x,y} v^2(b)$

$$
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$$
with \( \tilde{v}^2(b) := r^2(s) \) if \( g^b = \tilde{g}^b = 0 \).

Next we need to use large quantities \( Q_1 > 1 \) and \( Q_2 > 1 \) that will be determined precisely later (in the next section) by the choice of \( \alpha, \epsilon, \tau, \hat{\epsilon} \) and \( \check{\epsilon} \). Define \( L := Q_1 Q_2 \) and define \( K := L^{\left| \mathcal{N} \right|} \).

Define the function \( h : [1, \infty) \rightarrow [1, K] \) by \( h(r) = \min\{r, K\} \). Order the members \( \{s_1, \ldots, s_m\} \) of \( \mathcal{N} \) with \( \tilde{a}_{x,y}(s_1) \leq \tilde{a}_{x,y}(s_2) \leq \ldots \tilde{a}_{x,y}(s_m) \). Define \( L := Q_1 Q_2 \) and define \( K := L^{\left| \mathcal{N} \right|} \).

Define the function \( h : [1, \infty) \rightarrow [1, K] \) by \( h(r) = \min\{r, K\} \). Order the members \( \{s_1, \ldots, s_m\} \) of \( \mathcal{N} \) with \( \tilde{a}_{x,y}(s_1) \leq \tilde{a}_{x,y}(s_2) \leq \ldots \tilde{a}_{x,y}(s_m) \). Define \( \tilde{w}_{x,y}(s_k) = \prod_{j=k}^{m-1} h(\tilde{a}_{x,y}(s_{j+1})/\tilde{a}_{x,y}(s_j)) \).

For any move \( b \) at a state \( s \in \mathcal{N} \) define \( \bar{g}^b_{x,y} \) to satisfy
\[
(1 - \tilde{g}^b_{x,y}) = (1 - \bar{g}^b_{x,y})(1 - g^b_{x,y}). \tag{7}
\]

If \( \bar{g}^b_{x,y} = 1 \), then \( \bar{g}^b_{x,y} = 1 \) as well. Note that
\[
g^b v^2(b) + (1 - g^b)\bar{g}^b r^2(s) = g^b v^2(b) + (\bar{g}^b - g^b)r^2(s) = \bar{g}^b v^2(b). \tag{8}
\]

For every \( s \in \mathcal{N} \) and \( h \in \hat{\mathcal{H}} \) denote \( N^s(h) = \#\{n \in \mathcal{N} | s_n = s\} \in \mathbb{N} \cup \infty \). For \( 1 \leq i \leq N^s(h) \) let \( n^*_i(h) \) be the stage with the \( i \)th occurrence of the state \( s \) in \( h \). If the initial state of \( h \) is \( s \), then \( n^*_1 = 0 \) and \( N^s(h) \geq 1 \).

Define the discounted evaluation of a move \( b \) at a state \( s \in \mathcal{N} \) according to
\[
\xi^b_{x,y} = E^b_{x,y} r^2(h) \left( \sum_{i=1}^{N^s(h)-1} \bar{g}^{n^*_i(h)}(1 - \frac{\delta}{\bar{w}_{x,y}(s)})^{i-1} \prod_{k=1}^{i-1} (1 - \bar{g}^{n^*_k(h)}) + \right)
\]
\[
(1 - \frac{\delta}{\bar{w}_{x,y}(s)})^{N^s(h)-1} \prod_{k=1}^{N^s(h)-1} (1 - \bar{g}^{n^*_k(h)}) \right), \tag{9}
\]
where \( E^b_{x,y} \) stands for the expectation over all infinite histories \( h \in \hat{\mathcal{H}} \) with initial state \( s_0 = s \), assuming that Player Two plays the action \( b \) at stage 0, the first stage, and afterwards follows \( y \), whereas Player One follows \( x \) always.

**Lemma 5.1** The function \( \xi^b_{x,y} \) obeys the properties
\[
\xi^b_{x,y} = \bar{g}^b_{x,y} \tilde{v}^2_{x,y}(b) + (1 - \frac{\delta}{\bar{w}_{x,y}(s)}) (1 - \bar{g}^b_{x,y}) \xi_{x,y}(s) \tag{10}
\]
\[ r_{x,y}^2(s) = \xi_{x,y}(s) \left( 1 + \frac{\delta(1 - \tilde{a}_{x,y}(s))}{\tilde{w}_{x,y}(s)\tilde{a}_{x,y}(s)} \right) \]  

(11)

where \( \xi_{x,y}(s) = \sum_{b \in A^x} y^b \xi_{x,y}^b. \)

**Proof:**

We now verify that \( \xi \) satisfies (10) and (11). Separate the summation in (9) into three parts.

- All histories such that \( N^s(h) = 1 \). The probability of this event is \( g^b \), and the conditional expectation is \( v^2(b) \).

- All histories such that \( N^s(h) > 1 \) and \( i = 1 \). The probability of this event is \( 1 - g^b \), and the conditional expectation is \( \tilde{g}^b \gamma^2(s) \).

- All histories such that \( N^s(h) > 1 \) and \( i > 1 \). The probability of this event is \( 1 - g^b \). Factor out one power of \( (1 - g^b)(1 - \delta \tilde{w}) \); the conditional expectation is \( (1 - g^b)(1 - \delta \tilde{w}) \xi(s) \). By (7) this part contributes \( (1 - g^b)(1 - \delta \tilde{w}) \xi(s) \) to the sum.

Putting together the three parts, with (8) connecting the first two parts, we get (10). For equation (11) we use (10) and take the expectation with respect to the moves.

Notice that formula (11) is a slight variation of the standard relationship between discounted and undiscounted evaluations. \( \xi \) will serve as the auxiliary discounted payoff evaluation of Player Two. Note that \( r_{x,y}^2(s) \geq \xi_{x,y}(s) \forall s \in \mathcal{N} \). Define \( \tilde{\xi}_{x,y}(s) \) to be maximal value \( \max_{b \in A^x} \xi_{x,y}^b \).

**Lemma 5.2:** For every \( s, t \in \mathcal{N}, \gamma > 0 \), and \( (\delta, x, y) \in (0,1] \times X \times Y \)

- \( \tilde{a}_{x,y}(t) \leq K\tilde{a}_{x,y}(s) \) implies that \( \tilde{w}_{x,y}(t)\tilde{a}_{x,y}(t) \leq \tilde{w}_{x,y}(s)\tilde{a}_{x,y}(s) \),

- \( \tilde{w}_{x,y}(s)\tilde{a}_{x,y}(s) \leq \tilde{w}_{x,y}(t)\tilde{a}_{x,y}(t) \) and \( r_{x,y}^2(s) \leq r_{x,y}^2(t) + \gamma \) imply that \( \xi_{x,y}(s) \leq \xi_{x,y}(t) + \gamma + \delta \).

**Proof:** The first part follows directly from the definition of \( \tilde{w} \). For the second part, note that for every \( r, \tilde{w}, a > 0 \) and \( 0 < \delta < 1 \)

\[
\frac{r\tilde{w}a}{\tilde{w}a + \delta(1 - a)} - \frac{r\tilde{w}a}{\tilde{w}a + \delta} = \frac{r\tilde{w}a^2\delta}{(\tilde{w}a + \delta)(\tilde{w}a + \delta - \delta a)} \leq \frac{r\tilde{w}a^2\delta^2}{\tilde{w}^2a^2} = \frac{r\delta}{\tilde{w}}.
\]
Moreover, \( \frac{\min \{ \hat{w}, 1 \} \delta}{\hat{w} + \delta} \) is an increasing function in \( \hat{w}a \). Given \( \hat{w} \geq 1 \), from the above we have that \( r^2(s) \) and \( \hat{w}(s)a(s) \) determine \( \xi(s) \) except for a quantity of no more than \( \delta \). \( \square \)

### 5.2 The Best Reply Correspondence

For every state \( s \in \mathcal{N} \) define

\[
B^*_{\delta,1}(x, y) = \arg\max_{a \in A_1^s} u^1_{x,y}(a)
\]

\[
B^*_{\delta,2}(x, y) = \arg\max_{b \in A_2^s} \xi^b_{x,y} \text{ if } \xi_{x,y}(s) > j^\alpha_x(s)
\]

\[
B^2_{\delta,2}(x, y) = J^\alpha_x(s) \cup \arg\max_{b \in A_2^s} \xi^b_{x,y} \text{ if } \xi_{x,y}(s) = j^\alpha_x(s).
\]

\[
B^*_{\delta,2}(x, y) = J^\alpha_x(s) \text{ if } \xi_{x,y}(s) < j^\alpha_x(s).
\]

Player One maximizes her un-discounted payoff, while Player Two maximizes his auxiliary payoff, given that it is not too small.

Let the correspondences \( \overline{B}^\delta_{\delta,1} \) and \( \overline{B}^\delta_{\delta,2} \) be those defined by the closure of the graphs of the correspondences \( B^*_{\delta,1} \) and \( B^*_{\delta,2} \) in \( (X \times Y) \times A_1^s \) and \( (X \times Y) \times A_2^s \), respectively. Define \( \text{conv}(\overline{B}^\delta_{\delta,1}) \) and \( \text{conv}(\overline{B}^\delta_{\delta,2}) \) to be the correspondences with graphs in \( (X \times Y) \times X^s \) and \( (X \times Y) \times Y^s \), respectively, such that \( z \in \text{conv}(\overline{B}^\delta_{\delta,1}(x, y)) \) if and only if \( \{a \in A_1^s \mid z_a > 0\} \) is a subset of \( \overline{B}^\delta_{\delta,1}(x, y) \) and \( z \in \text{conv}(\overline{B}^\delta_{\delta,2}(x, y)) \) if and only if \( \{b \in A_2^s \mid z_b > 0\} \) is a subset of \( \overline{B}^\delta_{\delta,2}(x, y) \). Define the correspondences \( B_{\delta,1} \) from \( X \times Y \) to \( X \) so that \((x, y)\) in the domain corresponds to the sets \( B^\delta_{\delta,1}(x, y) \) in the range, and likewise define the correspondence \( B_{\delta,2} \) from \( X \times Y \) to \( Y \). We define the correspondence \( F_\delta : X \times Y \rightarrow X \times Y \) by \( F_\delta(x, y) = (B_{\delta,1}(x, y), B_{\delta,2}(x, y)) \). By Kakutani’s fixed point theorem for every \( \delta > 0 \) the correspondence \( F_\delta \) has a fixed point.

### 5.3 Two Lemmas on Fixed Points

We assume in the rest of the section that \((x, y)\) is a fixed point for \( F_\delta \). We prove Lemmata 5.4 and 5.5, described in the introduction.

**Remark 5.3:** Since the jump correspondence is used before \( \xi \) gets close to 0, any fixed point \((x, y)\) of \( F_\delta \) is absorbing. This implies that \( r^2_{x,y}(s) \geq j^\alpha_x(s) \) \( \forall s \in \mathcal{N} \). Indeed, suppose for the sake of contradiction that \( r^2_{x,y}(s) < j^\alpha_x(s) \). Denote by \( e \) the stopping time that is defined by the first stage in which the game leaves the set \( \{u \mid \xi_{x,y}(u) < j^\alpha_x(u)\} \). Recall from Section 2 that \( j^\alpha_x \) is
a sub-martingale. Since for every absorbing state \( s \in A \) \( \xi(s) = j^\alpha(s) = r^2(s) \) we have \( j^\alpha_x(s) \leq E_j^\alpha_x(s_e) \leq E \xi_{x,y}(s_e) \leq E r^2_{x,y}(s_e) = r^2_{x,y}(s) \), as desired.

**Lemma 5.4** If \( \tau \leq \omega \alpha/4 \) then there is a choice for \( L^* > 1 \) and \( \delta^* > 0 \) such that if \( L \geq L^* \) and \( 0 < \delta \leq \delta^* \) and \( (x, y) \) is a fixed point of \( F_\delta \) then

1) \( \xi_{x,y}(s) \geq j^\alpha_{x,y}(s) \) for all \( s \in \mathcal{N} \),
2) if the jump correspondence is used at \( s \) then \( \xi_{x,y}(s) \leq r^2_{x,y}(s) - 3\tau \)
3) for any action \( b \) from \( J^\alpha_x(s) \) used in \( y^s \ g^b_{x,y} < \tau \), and
4) the overall probability that Player Two plays an action from \( J^\alpha_x(s) \) at any \( s \in \mathcal{N} \) is at most \( \omega \alpha/20 \).

**Proof:** Let \( L^* = \frac{100|\mathcal{N}|}{\omega^2 \alpha^2} \) and \( \delta^* = \tau \alpha^3 \omega^3/(300|\mathcal{N}|) \). Choose \( t \) to be a member of \( \mathcal{N} \) with the largest difference \( j^\alpha_x(t) - \xi(t) \), and we must presume that this difference is non-negative. We will show that this difference can be no more than 0 and that the frequency devoted to the jump correspondence at any such state can be no more than \( \alpha \omega/20 \).

We presume for the sake of contradiction that the frequency devoted to the jump correspondence at \( t \) is at least \( \alpha \omega/20 \). Since \( r^2 \geq j^\alpha_x \) the expected value of the jump function \( j^\alpha_x \) at the states reached on the next stage after \( t \) using the jump correspondence \( J^\alpha_x \) is at least \( \alpha \omega \) more than \( j^\alpha_x(t) \), we must assume for any move from \( J^\alpha_x(t) \) that there is at least one state \( u \) reached by this move with a probability of at least \( \frac{\alpha^2 \omega^2}{40|\mathcal{N}|} \) such that \( j^\alpha_x(u) \geq j^\alpha_x(t) + \alpha \omega/2 \), necessarily with \( \text{esc}(u, t) \leq \alpha \omega/4 \). (If \( \text{esc}(u, t) > \alpha \omega/4 \) then \( a(t) \geq \alpha^3 \omega^3/(160|\mathcal{N}|) \) and by (11) and the size of \( \delta^* \) we have made \( \xi(t) \) too close to \( r^2(t) \) contradicting \( j_x(t) \leq r^2(t) - \alpha \omega/2 \), – which must follow by Remark 5.3 since otherwise any move from the jump correspondence would be evaluated in an undiscounted way strictly above the level \( j^\alpha_x(t) \).) By the definition of \( \bar{\omega} \), the size of \( L^* \) and (3) we have \( \bar{\omega}(u) \bar{a}(u) \). By \( \text{esc}(u, t) \leq \alpha \omega/4 \) it follows that \( |r^2(t) - r^2(u)| \leq \alpha \omega/4 \). But by Lemma 5.2 we have \( \xi(t) \geq \xi(u) - \delta - \alpha \omega/4 \). With the size of \( \delta^* \) this contradicts \( j^\alpha_x(u) \geq j^\alpha_x(t) + \alpha \omega/2 \) and the choice of \( t \).

Next, suppose for the sake of contradiction that \( J^\alpha_x \) is used at \( s \) and \( g^b \geq \tau \) for some move \( b \in J^\alpha_x(s) \). Indeed, \( g^b \geq \tau \) implies that \( \tilde{g}^b = 1 \). In particular, using Remark 5.3, \( \xi^b_{x,y} = w_{x,y}(b) \geq \sum_t p(t|s; x, b) j^\alpha_x(t) \geq j^\alpha_x(s) + \omega a \). Thus, for every \( b' \in B^2_\delta(x, y) \) that maximizes \( \xi \), \( \xi^b_{x,y} \geq \xi^b_{x,y} \geq r^2_{x,y}(s) \geq j_x(s) + \omega a/2 \). Since the overall probability to play actions from the jump correspondence is smaller than \( \omega \alpha/20 \), this contradicts the assumption \( \xi(s) \leq j^\alpha_x(s) \).
Now we presume for the sake of contradiction that \( \xi(s) \geq r^2(s) - 3\bar{\tau} \) and the \( J^\alpha_x \) correspondence is used at \( s \). Since we must assume that \( \xi(s) = \bar{j}^\alpha_x(s) \), we have an increase in the value of \( r^2 \) of at least \( \omega\alpha - 3\bar{\tau} \) from a move in \( J^\alpha_x \). By the dominance of \( \omega\alpha \) over \( 4\bar{\tau} \), we must conclude that \( g^b > \bar{\tau} \), a contradiction. □

Lemma 5.4 is the most problematic aspect of extending this proof to the case of finitely many positions. Any identification of infinitely many states as a single state may be meaningless if the states reached from it are not also identified. A more flexible definition of the discounted evaluation may be necessary. For example, at a state \( s \) one could discount future visits to other states \( t \) according to the difference between Player Two’s undiscounted expected payoffs from these two states.

The following lemma claims that if the auxiliary payoff is too far from the real payoff and the action causes absorption with small probability, then this probability is very small. This radical discontinuity is the key argument to our whole approach.

**Lemma 5.5** For \( L, \alpha, \tau \) and \( \delta \) satisfying the conditions of Lemma 5.4 and \( (x, y) \) a fixed point of \( F_\delta \) if \( \xi(s) \leq r^2(s) - 2\bar{\tau} \) and \( g^b \leq \bar{\tau} \) then \( g^b \leq 1.1\delta\xi(s)/\bar{w}(s) \) and \( g^b \leq 2.3\bar{\tau}a(s) \leq 2.3a(s) \).

**Proof:** First we claim that \( \bar{\xi}(s) - \xi(s) < \frac{\delta\alpha\omega}{19a(s)}\xi(s) \).

If the jump correspondence at \( s \) is used and \( b \) is such a move, since \( g^b \leq \bar{\tau} \) (from Lemma 5.4) it follows that \( g^b = g^b/\bar{\tau} \). Hence from (6) we have

\[
\bar{v}^2(b) = (1 - \bar{\tau})r^2(s) + \bar{\tau}\bar{v}^2(b) \geq r^2(s) - \bar{\tau} \geq \xi(s) + \bar{\tau}.
\]

Moreover, from (10) and (12) we have

\[
\xi^b \geq \xi(s) + \bar{g}^b(\bar{v}^b - \xi(s)) - \delta\xi(s)/\bar{w}(s) \geq \xi(s)(1 - \delta/\bar{w}(s)),
\]

and by Lemma 5.4, since \( \xi(s) \) is the average of \( \bar{\xi}(s) \) and such \( \xi^b \), we have \((1 - \alpha\omega/20)(\bar{\xi}(s) - \xi(s)) \leq \frac{\delta\alpha\omega\xi(s)}{20\bar{w}(s)} \), so the claim follows.

Considering now any move \( b \in A^*_x \) that is used with \( g^b \leq \bar{\tau} \) and looking again at formula (10) we have \( \bar{\xi}(s) \geq \xi^b \geq \bar{g}^b(\bar{v}^b - \xi(s)) + (1 - \delta/\bar{w}(s))\xi(s) \) and hence \( \bar{g}^b(\bar{v}^b - \xi(s)) \leq 1.1\delta\xi(s)/\bar{w}(s) \), since by the above claim \( \xi^b(s) - \xi(s) \) is small compared to \( \frac{\delta}{\bar{w}(s)}\xi(s) \). First consider the consequence of \( \bar{v}^b - \xi(s) \geq \bar{\tau} \), namely \( g^b = \bar{\tau}g^b \leq 1.1\delta\xi(s)/\bar{w}(s) \). Second, consider \( g^b \leq
Lemma 6.1 For every two distinct non-absorbing states \( s, t \) with \( \text{esc}(t, s) \leq \gamma < 1 \) in an absorbing time homogeneous Markov chain \( P^t(t, s)\mu(s, t) \) does not differ from \( a(t) \) by more than a factor of \( 2\gamma \), \( \text{esc}(t, s)/\mu(t, s) \) is within a factor of \( 3\gamma \) to the ratio that, starting at \( t \) or \( s \), the last visit before absorption was at \( t \) rather than at \( s \). Furthermore, with or without the assumption that the Markov chain is absorbing and with a start at either \( s \) or \( t \), the ratio of the expected number of visits to \( s \) to those at \( t \) is at least \( 1 - 4\gamma \) times the ratio of \( P^t(t, s) \) to \( P^n(s, t) \).

Proof: The first two claims follow directly from the formulas (1) and (2). The third claim follows from the first claim if the Markov chain is absorbing. Otherwise we recognize in \( 1/P^t(t, s) \) the expected number of visits to \( t \) before reaching \( s \).

Remark 6.2 At a fixed point of \( F_\delta \) satisfying the properties of Lemmata 5.4 and 5.5, if \( \xi(s) \leq r^2(s) - \bar{\tau} \) and \( b \) is a move at \( s \) with \( g^b \geq \bar{\tau} \) then \( w^2(b) = \bar{\tau}(s) \), which is by Lemma 5.5 also within \( \delta/20 \) of \( \xi(s) \).

Theorem 2: For any choice of positive \( \epsilon, \bar{\tau}, \hat{\epsilon}, \bar{\bar{\epsilon}} \) satisfying the inequalities stated in Theorem 1 all conditions of Theorem 1 are satisfied.

Proof: Because it is sufficient to demonstrate the conclusion of Theorem 1 with smaller choices for \( \bar{\tau}, \hat{\epsilon}, \bar{\bar{\epsilon}} \), we will assume without loss of generality that \( \alpha \) is small enough so that for every \( s \in \mathcal{S} \) \( c^\alpha(s) \) is within \( \epsilon/2 \) of the undiscounted zero-sum value \( c_2(s) \), as described in Section 2, and \( \bar{\tau} < \alpha \omega/4 \).

Define \( \beta := \frac{1}{2} \hat{\epsilon} \hat{\epsilon} |N| / (3|N||N|) \). We require that \( L := Q_1 Q_2 \) is large enough
to satisfy the conditions of Lemma 5.4 and also that
\[ Q_1 > 80|\mathcal{N}|^2 m^2 / (\rho \tilde{\varepsilon}^2 \tilde{\varepsilon}^2 \beta^2) \quad \text{and} \quad Q_2 > 80m|\mathcal{N}| / (\rho \tilde{\varepsilon} \tilde{\varepsilon}). \]

We begin with \( \delta \) sufficiently small, so that the condition of Lemma 5.4 holds. Next, we consider fixed points of \( F_\delta \) corresponding to decreasing \( \delta > 0 \) that have convergent subsequences for certain variables living in compact spaces – the stationary strategies in the space \( \mathcal{X} \times \mathcal{Y} \), the values \( \nu(a, b) \) for all pairs of moves at all states, the expected payoffs \( r^1(s), r^2(s) \), and the absorption rate \( a(s) \) for every \( s \in \mathcal{N} \), and the probabilities \( \text{esc}(t, s) \) for all pairs of states.

We define a move \( a \in A^1 \) or \( b \in A^2 \) to be a limit move if and only if the frequency of its use does not converge to zero as \( \delta \) goes to zero, and define \( \hat{q} \) to be the minimal positive limit value for a frequency of a limit move chosen by either player. We define the quantity \( \hat{\mu} \) to be the minimal positive limit value for \( \text{esc}(s, t) \), \( \hat{\nu} \) to be the minimal positive limit value for \( \nu(a, b) \), and \( \hat{a} \) the minimal positive limit value for \( a(s) \).

Next we must define the partition \( \mathcal{R} \) of a subset \( P \). Define a directed graph on the space \( \mathcal{N} \) by \( t \to s \) if and only if in the limit \( \text{esc}(t, s) \) approaches zero. The relation is transitive, but not necessarily symmetric. It has an additional property, that if \( t \to s_1 \) and \( t \to s_2 \) then either \( s_1 \to s_2 \) or \( s_2 \to s_1 \). This is easy to confirm, because if \( s_1 \) was not reached with probability approaching one on the way from \( t \) to \( s_2 \) then it must be reached with probability approaching one after the state \( s_2 \). Next define a relation \( \sim \) that is symmetric, transitive, and reflexive on an appropriate subset; \( s \sim t \) if and only if \( \mu(s, t) \) approaches zero, and \( s \sim s \) if and only if \( a(s) \) approaches zero. \( \sim \) defines a partition \( \mathcal{P} \) of a subset \( P' \) of \( \mathcal{N} \). Now we relate \( \to \) to \( \sim \). Define \( \mathcal{R} \) to be the subset of \( \mathcal{P} \) defined by \( A \in \mathcal{R} \subseteq \mathcal{P} \) if and only if \( u \in A \) and \( u \to s \) implies that \( s \in A \). Any state \( s \notin A \in \mathcal{R} \) such that \( \text{esc}(s, u) \) approaches zero for any (equivalently some) state \( u \in A \in \mathcal{R} \) is called a satelite of \( A \). Due to the above, a satelite of \( A \in \mathcal{R} \) cannot be a satelite of any other member of \( \mathcal{R} \) and every member of \( Q \in \mathcal{P} \) such that \( Q \) is not in \( \mathcal{R} \) must be a satelite of the same \( A \in \mathcal{R} \). We call an primitive exit \((a, b)\) from \( R \in \mathcal{R} \) to be a satelite exit if with certainty the exit results in motion that doesn’t leave \( R \) or its collection of satelites.

For every \( R \in \mathcal{R} \) we define the set \( B_R \) of Player Two moves in \( R \) to be \( B_R := \{ b \in A^2_2 \mid \text{for some limit move } a \in A^1 \ (a, b) \ \text{is an primitive exit from } R \text{ that is not a satelite exit} \} \).

If \( s \) is a satelite of \( R \in \mathcal{R} \) then in the limit the probability that the last
visit to the pair $s$ or any $u \in R$ was the state $s$ must go to zero. Therefore $\nu(a, b)$ approaches zero for any satellite exit $(a, b)$ at $u \in R \in \mathcal{R}$. These facts follow directly from Lemma 6.1 and $\text{esc}(s, u)/\mu(u, s)$ going to zero in the limit.

We show for every $R \in \mathcal{R}$ and pair $s, t \in R$ that the probability of using some exit in $E^{R_s}(R)$ before reaching $t$ from $s$ also approaches zero. First this holds for any non-satellite primitive exit from $R$, because the probability of reaching a non-satellite outside of $R$ would be at least $\rho$ and therefore the probability of absorbing before reaching $t$ must be in the limit at least the probability of using this exit times $\hat{\mu}\rho$. The same arguments hold for the use of any move in $B_R$, but with the quantity $\hat{\mu}\rho\hat{q}$ instead of $\hat{\mu}\rho$.

More difficult is to show that the above holds for any satellite exit $(a, b)$ at $u \in R$. Let $v$ be any satellite of $R$ reached with positive probability from this exit. Let $\pi$ be the probability of using $(a, b)$ before reaching $t$ from a start at $s$ and let $\theta$ be a bound on the probability of not reaching any member of $R$ from any other member of $R$ or from a satellite of $R$. Let $\hat{\gamma}$ be the probability of reaching $v$ from $s$ before reaching $t$, with $\hat{\gamma} \geq \pi\rho$. Going through the state $v$, the probability of reaching $t$ is at least $1 - \theta$ and the combined probability of reaching $t$ from $s$ is also at least $1 - \theta$. This means that the probability of reaching $t$ from $s$ conditioned on not going through $v$ is at least $1 - \frac{\theta}{(1-\gamma)}$. So conditioned on not arriving at $v$ before $t$ there is at most a $\frac{2\theta}{1-\gamma}$ probability of absorbing before getting back to $s$. In the limit $\frac{2\theta}{1-\gamma}$ cannot stay above 1, because $\theta$ goes to zero and in the limit $\hat{\gamma}$ cannot go above $1 - \hat{\mu}$. This means that eventually the probability of reaching $v$ from $s$ must be at least $\hat{\gamma}\sum_{i=0}^{\infty}(1 - \hat{\gamma})^i(1 - \frac{2\theta}{1-\gamma})^i = \hat{\gamma}\sum_{i=0}^{\infty}(1 - 2\theta - \hat{\gamma})^i = \frac{\hat{\gamma}}{2\theta+\hat{\gamma}}$. But this probability to reach $v$ from $s$ cannot go above $1 - \hat{\mu}$ in the limit, which is possible only if $\pi$ goes to zero as $\theta$ goes to zero also.

Define $\epsilon^*$ to be $(\hat{\nu}\hat{\mu}\hat{q}\hat{a}/K)^{3|\mathcal{N}|}$. We require of a fixed point of $F_\delta$ that the values for which we have convergent subsequences are within $\epsilon^*$ of their limit values. We require that the probability of using any exit before moving from any $s$ to $t$ for any pair $s, t \in R \in \mathcal{R}$ is no more than $\epsilon^*$ and for every $R \in \mathcal{R}$ that the sum of $\nu(a, b)$ over all the satellite exits $(a, b) \in E^{B_R}(R)$ is no more than $\epsilon^*$ (as demonstrated above). Furthermore we require that $\delta < (\epsilon^*)^2$. We let $(x, y)$ be a fixed point of $F_\delta$ satisfying these properties. If the stationary strategy is not specified, then $(x, y)$ is intended.

**Step 1:** For every $s \in R \in \mathcal{R}$ show that if $\frac{x y^2}{2|x| y}(s) \geq \beta$ then there
exists an $\tilde{e}$ simplification $\tilde{\pi}$ of $x$ such that $\frac{2}{\tilde{\pi}}(R) \geq 1 - 3|R|/Q_1\beta$ and for all $t \in R$ that $\frac{2}{\tilde{\pi}}(t)/\frac{2}{\tilde{\pi}}(s) \geq (1 - \frac{m}{2\rho Q_2} - \frac{4m|R|}{\rho Q_2})(\frac{2}{\tilde{\pi}}(t)/\frac{2}{\tilde{\pi}}(s))$:

For every $d \geq 1$ define

$$T_d = \{t \in R | \mu(s,t) \leq d\tilde{a}(s)\} \cup \{s\}.$$  

Denote $T = T_d$, where $d \in (1, L^{[N]-1})$ satisfies $T_{Ld} \setminus T_d = \emptyset$. Since $K = L^{[N]}$, for every $t \in T$ we have $\tilde{a}(t) \leq a(t)/\tau \leq \mu(s,t)/\tau \leq d\tilde{a}(s)/\tau \leq K\tilde{a}(s)$, and it follows that $\tilde{w}(t)\tilde{a}(t) \leq \tilde{w}(s)\tilde{a}(s)$. With $\xi(s) \leq r(s) - 2.5\tau$ we have by (11) that $\tilde{a}(s) \leq \delta/\tau$, $a(s) \leq \delta/\tau$, and $s, t \leq \delta K/\tau < \epsilon^*$, meaning that $T$ is a subset of $R$. Since $t \in T$ satisfies $|r^2(s) - r^2(t)| \leq \epsilon^*$ we have $\xi(t) \leq \xi(s) + (\epsilon^* + \delta)$.

Define a quantity

$$p^t = \begin{cases} 
\frac{a(t)}{Q_2}\mu(s,t) & t \in T \setminus \{s\} \\
1/Q_2 & t = s
\end{cases}$$

Define the stationary strategy $\pi$ by removing from $x$ all Player One moves at states $t \in T$ that are played with probability smaller than $p^t/\rho$, and normalize the remaining vector. This means that if $u$ is reached in one stage from $t \in T$ by $\pi$ and a Player Two move $b$, then $p(u|t; \pi, b) \geq p^t$.

We use critically from Lemma 6.1 that $a(u)/\mu(u,s)$ is approximately $P^u(u,s)$ (within a factor of $2\epsilon^*$) for any $u \in R$, so that from Lemma 3.2 and Lemma 6.1 with the change from $(x,y)$ to $(\pi,y)$ the ratio of visits at $t \in R$ to those at $s$ cannot increase by more than a factor of $\frac{8|T|m}{KQ_2}$. Furthermore, by the definition of $\pi$, $\tilde{v}, \delta \leq (\epsilon^*)^2$ and Lemma 5.5 there are no non-satellite exits performed inside of $T$ other than those generated by Player Two moves $b \in A_2$ with $w_{x,y}(b) = \xi_{x,y}(t)$. Combined with the fact that the absorption rate of any move $b$ with $g^b \geq 2.4\tilde{\pi}$ is altered by a factor or no more than $m/(2\rho Q_2)$ by the switch to $\pi$ and that $2\epsilon^*$ is greater than the probability that the last visit to $R$ was at a satelite exit, we have everything but the claim that there is only insignificant motion with $(\pi,y)$ toward absorption from states in $R$ outside of the set $T$.

We can break up the absorption from $R$ generated by the strategies $(x,y)$ in terms of where was the last visit in $R$. Let $t \in T$, $u \in R \setminus T$ and $b \in B$ be a move such that $p(u|t; \pi, b) > 0$, necessarily with $p(u|t; \pi, b) \geq p^t$. To complete the claim of Step 1 it suffices to show that $\frac{\text{OSC}_{x,y}(u,t)}{\mu_{x,y}(u,t)} \leq \frac{2.5}{Q_1}$ for every such $u \in R \setminus T$. 

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Case 1; \( u \in R \setminus T \) is reachable from \( T \) only by Player Two moves \( b \) with \( g^b \geq \frac{\epsilon}{2} \):

It follows immediately from the fact that Player Two has no more than \( m|R| \) moves in \( R \) that \( \frac{\text{esc}(u,t)}{\mu(u,t)} \) is smaller than \( 2.5/Q_1 \), since any such move doesn’t return to \( R \) with a probability of at least \( 2.5\epsilon \) and with at least \( 1-2\epsilon \) probability there is motion from \( u \) back to \( t \in T \).

Case 2; \( t \neq s \), and \( u \in R \setminus T \) is reachable by \((\bar{x}, y)\) from a \( t \in T \) by a move \( b \) of Player Two with \( g^b < \frac{\epsilon}{2} \):

By Lemma 5.5 we have

\[
p_t \leq \frac{a(s)}{4|N|} \leq 2.3a(s).
\]

Case 3; \( u \in R \setminus T \) is reachable by \((\bar{x}, y)\) from \( s \) by a move \( b \) of Player Two with \( g^b < \frac{\epsilon}{2} \):

We have \( p^* = 1/Q_2 \), \( p^* \leq 2.3a(s) \leq \frac{\epsilon}{2} \) and

\[
\frac{\text{esc}(u,s)}{\mu(u,s)} \leq \frac{2.3a(s)Q_2}{L} \leq 2.3a(s)Q_2 \leq 2.3a(s)L \frac{Q_1}{Q_1}.
\]

In all arguments that follow concerning members of a set \( T \) as created above, for convenience we will write \( z^\gamma \) or \( B^\gamma \) instead of \( z^\gamma \) or \( B^\gamma \) for \( \gamma > \frac{\epsilon}{2} \). By Lemma 5.5 there will be no difference in these expressions.

Step 2; For any choice of \( s \in R \) from Step 1 there is an \( \frac{\epsilon}{2} \) simplification \( \bar{y} \) of Player Two’s strategy \( y \) such that together with \( \bar{x} \) the state \( s \) and all states \( t \in T \) with \( \frac{\text{esc}(x,y)(t)}{\mu(x,y)(s)} \geq \frac{\epsilon}{2} \) \( \frac{z^\gamma(s)}{4|N|} \) are reached by \((\bar{x}, \bar{y})\) from all of \( R \), and furthermore from inside of \( T \) no state outside of \( T \) is reached:

We define \( \bar{x} \) for all \( t \in T \) by removing from \( y \) all moves made by Player Two with a frequency of \( L/(L-1)Q_1 \) or less, followed by normalization.

Any \( t \in T \) that satisfies \( \frac{\text{esc}(x,y)(t)}{\mu(x,y)(s)} \geq \frac{\epsilon}{2} \) \( \frac{z^\gamma(s)}{4|N|} \) by Step 1 also satisfies \( P_{\bar{x},y}(s,t) \geq \frac{\epsilon}{2} \) \( \frac{z^\gamma(s)}{4|N|} \) and \( z^\gamma(t) \geq \frac{\epsilon}{2} \) \( \frac{z^\gamma(s)}{4|N|} \). Notice that this last condition
is satisfied by the state $t = s$. For any $t \in T$ with $P^w_{x,y}(s, t) \geq \frac{\bar{\tau}(\xi, \beta)}{4.5|N|} P^w_{x,y}(s, t)$ and $\bar{z}_{x,y}(t) \geq \frac{\bar{\tau}(\xi, \beta)}{4.5|N|}$ to show that $t$ is reached from all of $T$ with $(\bar{x}, \bar{y})$ by Lemma 3.4 it suffices to show that for any $w \in T$ and any $t \in T$ satisfying $\bar{z}_{x,y}(t) \geq \frac{\bar{\tau}(\xi, \beta)}{4.5|N|}$, including $s = t$, we have that the change from $w^y$ to $\bar{w}^w$ does not reduce $P^w_{x,y}(w, t)$ by more than a factor of $\frac{\bar{\tau}(\xi, \beta)}{12|N|^2}$.

If $b \in A^w_2$ is a Player Two move with $g^b \geq \bar{\tau}$, removing $b$ to form $\bar{w}^w$ from $y^w$ cannot reduce $P^w_{x,y}(w, t)$ by anything more than a factor of $\frac{\bar{\tau}(\xi, \beta)}{12|N|^2}$. Assuming that $g^b < \bar{\tau}$ and removing $b$ to make $\bar{w}^w$ removes at least $\frac{\bar{\tau}(\xi, \beta)}{12|N|^2}$ of the motion $P^w_{x,y}(w, t)$ we would have from Lemma 5.5 that $2.3 a(w) L/Q_1 (L-1) \geq g^b_x y^b \geq \frac{\bar{\tau}(\xi, \beta)}{12|N|^2} \bar{z}_{x,y}(t)$. This is a contradiction to the definition of $Q_1$.

Second, we show that, starting at $s$, motion according to $(\bar{x}, \bar{y})$ never leaves the set $T$. Let us assume that $u$ is a state not in $T$ reached by a move $b$ of Player Two from any $t \in T$ played against $x$ and given positive frequency by $y$. We need to show that $b$ is not used in $y$. If $t \neq s$ then by formula (3) $\mu(t, u) \leq \frac{\alpha(t)}{\bar{\tau}^2 y_b} = \frac{\mu(s, t) Q_2}{y_b}$. In particular, by the definition of $T$ and since $\mu$ is a metric,

$$y_b \leq \frac{\mu(s, t) Q_2}{\mu(t, u)} \leq \frac{\mu(s, t) Q_2}{\mu(s, u)} \mu(t, u) \leq \frac{Q_2}{L-1} = \frac{L}{(L-1)Q_1}.$$  

And if $b$ is a move at the state $s$ then also by the definition of $T$ and (3)

$$y_b \leq \frac{\alpha(s) Q_2}{\mu(s, u)} \leq \frac{Q_2 \alpha(s)}{L \alpha(s)} \leq \frac{1}{Q_1}.$$  

Therefore we conclude that $(\bar{x}, \bar{y})$ defines one ergodic set $D \subseteq T$ that includes $s$ and all states $u \in R$ satisfying $\bar{z}_{x,y}(u) \geq \frac{\bar{\tau}(\xi, \beta)}{4|N|}$.

Step 3; Show that there is a proper choice of $s$ from Steps 1 and 2 with a subset $V_{R}$ of Player Two moves satisfying the conditions of Theorem 1, namely that these moves belong to a subset $F$ containing $s$ and inside of the ergodic set $D$ such that $\bar{v}(t) \bar{a}(t)$ is a constant for all $t \in F$ and there is a distribution on $V_{R}$ such that used against $x$ gives an expected payoff to Player One within $\epsilon$ of $r^1(s)$:

Define $U := \{ t \in R \mid \xi(t) \leq r^2(t) - 2.4 \bar{\tau} \}$ and define $\bar{U} := \{ t \in R \mid \xi(t) \leq r^2(t) - 2.5 \bar{\tau} \} \cap \{ t \in R \mid \bar{z}_{x,y}(t) \geq \beta \}$. We create a partition $\{U_1, \ldots, U_k\}$ of
the members of $U$ in increasing values of $\tilde{w}a$, meaning that $s$ and $t$ belong to the same member of $U_i$ if and only if $\tilde{w}(s)a(s) = \tilde{w}(t)a(t)$. For any state $s$ in $\tilde{U}$ we consider the sets $T(s)$ and $D(s)$ and the strategies $\pi(s), \bar{\pi}(s) \in X \times Y$ as created above in Step 1 and Step 2.

For the sake of contradiction we suppose that there is no $s \in \tilde{U} \cap U_i$ and $b \in B_{x,y}(t)$ with $t \in D(s) \cap U_i$ such that $|v_{\pi(s),y}(b) - r^1(s)| \leq \tilde{\epsilon}$ and there is no pair of Player Two moves $b, b' \in B_{x,y}(R)$ with both $b$ and $b'$ belonging to the set $D(s) \cap U_i$ with $v_{\pi(s),y}(b)$ and $v_{\pi(s),y}(b')$ on different sides of $r^1(s)$.

For every $s \in \tilde{U}$ and $t \in U \cap D(s)$ with some move in $B_{x,y}(t)$ used in $y_t$ let $v^1_s(t) = \sum_{b \in B_{x,y}(t)} v^1_{\pi(s),y}(b)\nu^b_{\pi(s),y} / \sum_{b \in B_{x,y}(t)} \nu^b_{\pi(s),y}$, the average Player One payoff resulting from these moves at $t$. For every $1 \leq i \leq k$ let $p(i) := \sum_{j<i} |U_j|.$

We claim that our above assumption implies that $z_{x,y}^\pi(s) \leq 3p(i) \beta/\tilde{\epsilon}p(i)$ for every $s \in U_i \cap \tilde{U}$.

We prove the above claim by induction on $i$. Let $s$ be any member of $U_i \cap \tilde{U}$, and we assume that $|v^1_\pi(s) - r^1_{\pi(s),y}(s)| = |v^1_\pi(s) - r^1_{x,y}(s)| \geq \tilde{\epsilon}$. From Part 1 and Part 2 we know that the importance with respect to $(\pi(s), y)$ from exits outside of $B_{x,y}(D(s))$ does not exceed $\bar{\pi} \tilde{\epsilon}/3$ times the importance of the $B_{x,y}(s)$ moves. Since all the $v^1_s(t)$ with $t \in D(s) \cap U_i$ are on the same side of $r^1(s)$ as $v^1_\pi(s)$, we are left only with the $B_{x,y}$ moves from $\cup_{k<i} U_k \cap D(s)$ to counter-balance the $v^1_s$ to make $r^1_{\pi(s),y} = r^1_{x,y}$. We can assume now that $i > 1$, since otherwise we would have to conclude that $|v^1_\pi(s) - r^1_{x,y}(s)| \geq \tilde{\epsilon}$ is impossible. By the induction hypothesis the sum of all the $z_{x,y}^\pi(u)$ over the set $\cup_{k<i} U_k$ does not exceed $\sum_{k<i} |U_k| 3p(k) \beta/\tilde{\epsilon}p(k) \leq \frac{2}{3} 3p(i) \beta/\tilde{\epsilon}p(i-1)$. By the fact that our simplications $\bar{\pi}(s)$ hardly influence the expected payoffs from moves with an absorption rate of at least $2\bar{\pi}$ and by the statement of Step 1, in order to maintain $|v^1_\pi(s) - r^1_{x,y}(s)| \geq \tilde{\epsilon}$ we must assume that $\tilde{\epsilon} z_{x,y}^\pi(s) \leq \frac{2}{3} 3p(i) \beta/\tilde{\epsilon}p(i-1)$, and this concludes the proof of our claim.

With the definition of $\beta$ we conclude that $z_{x,y}^2 \bar{\pi}(s) < \tilde{\epsilon}/|R|$ for every $s \in R$, and this means that $R$ could not have been chosen for polarization, a contradiction.

With the appropriate $s \in R$ chosen, we have $D_R := D(s)$, $x_C$ defined from $\bar{\pi}(s)$ and $y_C$ defined from $\bar{\pi}(s)$ so that changes are made only inside of $D_R$, and the exits $V_R$ and their distribution as determined by $y_D$ come from the above argument.
Step 4; show that the moves $B_R$ satisfy the requirements of Theorem 1:

The easiest way to prove that $|r_{x,y}(s_R) - r_{x,y'}(s_R)| \leq \epsilon$ is to return part of the way back to the space $S_2$! We let $S_2$ be the space generated by the almost trivial partition $P := \{R\} \cup \{s \mid s \not\in R\}$. With $\mu^1$ the harmonic function on $S_2$ induced by $\mu^1$ on the absorbing states, by Lemma 3.5 $\mu^1(s_R)$ and $r^1(s_R)$ differ by at most $4\epsilon^*$. Let $\nu_2$ be the corresponding measure of the importance of the exits.

Define a move $a \in A_1^s$ of Player One in the set $R$ to be a principle move if $a$ is not a limit move and if there is a $b \in B_2^s$ such that $(a,b)$ is an exit with $\nu_{x,y}(a,b) \geq \hat{\nu} - \epsilon^*$. We claim that the combination $(a,b)$ of a move of $B_R$ with a principle move of Player One must yield $\nu(a,b) \leq \epsilon^*$. Once this is established from the definition of $B_R$ we need only to break down the sum of the $\nu^1(a,b)\nu_2(a,b)$ over all exits $(a,b)$ with $\nu(a,b) \geq \hat{\nu} - \epsilon^*$ and apply Lemma 3.7 to conclude that $r_{x,y}(s_R)$ is within $20\|me^*/\hat{\nu}\|$ of $r_{x,y}(s_R)$, that is much closer than we need it. Suppose for the sake of contradiction that for some principle $a$ and some $b \in B_R$ that $\nu(a,b) \geq \hat{\nu} - \epsilon^*$. Assuming that the moves take place at $t$, we have from the definition of $B_R$ that $a(t) \geq y_b(\hat{\nu} - \epsilon^*)(\hat{\mu} - \epsilon^*)$. Furthermore by definition we have $\nu(a,b) \leq x_a y_b / a(t)$ and by assumption $x_a \leq \epsilon^*$. These four inequalities are contradictory.

We show that $b \in B_R \cap A_2^s$ with $t \in P$ implies $v^2(b) < r^2(t)$. If $\xi(t) \leq r^2(t) - 2\tau$ then it follows from Lemma 5.5. If $\xi(t) > r^2(t) - 2\tau$ then by Lemma 5.4 all moves have the same auxiliary value $\xi(t) = \overline{\xi}(t)$; it follows from the smallness of $\delta \leq (\epsilon^*)^2$ and formulas (10) and (11) that if $v^2(b) \geq r^2(t)$ then the repeated use of $b$ would result in a higher evaluation for $\xi(t)$ because an undiscounted value of at least $r^2(t)$ would be obtained but at much higher auxillary absorbing rate.

Step 5; show $z^{2,5\tau}(s) < \epsilon$ for any state $s$ that is not in $P$ or is a satelite of some $R \in \mathcal{R}$:

If $s$ is not a satelite and not in $P$ then due to the very small size of $\delta$ we have from (11) that $\xi(s)$ is within $\epsilon$ of $r^2(s)$, implying that no move $b$ used at $s$ could satisfy $w^2(b) < r^2(s) - 2\tau$. For a satelite $s$ of $R$ we suppose that $b \in A_2^s$ is a Player Two move at $s$ with $g^b \geq 2.5\tau$. Such moves have at least a $2.4\tau$ probability of never returning to the set $R$. The probability of using
such a move before reaching \( R \) must be no more than \( \epsilon^*/(2.4\tau) \), and thus the total probability that it is used cannot exceed \( \epsilon^*/(2.4\tau(\hat{\mu} - \epsilon^*)) \). q.e.d.

### 7 Signaling

In this section we show that there are approximate equilibria without an assumption that Player One can send signals independent of the transitions. The problem concerns the consequences to the players of any moves that would be used by Player One as a transition dependent signal. For example, a move of Player One that brings the play outside of the set \( D_R \) may fail to be useful to signal her desire for Player Two to use a move in \( V_R \), because outside of \( D_R \) the jump function for Player Two may exceed greatly his expected payoff from the moves in \( V_R \).

The natural solution is for Player One to have a move inside of \( D_R \) that is not used in \( x_C \) whose use means that the moves \( V_R \) of Player Two will not be used, and after a certain quantity of visits to some state in \( D_R \) it will be understood mutually that Player Two must use a move in \( V_R \). A problem arises, however, if every such move results in a positive probability of leaving the set \( R \).

With regard to the next two theorems, we assume the statement and proof of Theorem 2, which means also that we assume that all the conditions of Theorem 1 are satisfied. We will add new conditions to those of Theorem 1 and make some minor changes to the proof of Theorem 1. The definition of \( S_2 \) remains, along with its Markov chain transitions, including the \( p^*_R \) and \( p_R \).

The changes begin with the definition of the parts \( q^R_{dR} \) and \( q^R_{tR} \) and therefore everything that follows in the proof of Theorem 1 will be altered as well, including the introduction of new situations.

\textbf{Theorem 1':} Assume the following property for every \( R \in \mathcal{R} \): if every move \( a \in A^1_t \) in \( D_R \) removed to make \( x_C \) from \( x \) formed an exit against some Player Two move used in \( y_C \), then there exists a set \( A^1_R \) of Player One principle moves in \( D_R \) such that

1) the sum of \( \nu_{x,y}(a,b) \) for all \( R \) exits \((a,b)\) performed outside of \( D_R \) does not exceed \( \epsilon \bar{\epsilon} \bar{\tau} \bar{\beta} / 3 \),

2) for every principle move \( a \in A^1_R \) of Player One used at \( t \in D_R \) with \( \nu^a_{x,y} \geq \beta \bar{\epsilon} \tau / (3|N|m) \) we have \( \sum b \) used in \( y_C \) \( \nu^a_{x,y}(a,b) \geq (1 - \tau \bar{\epsilon} \bar{\beta}) \nu^a_{x,y} \)
and therefore also $|v^1_{x_C,y_C}(a) - v^1_{x,y}(t)| \leq \hat{\epsilon}$.

**Conclusion:** Without any assumption on Player One’s ability to signal independent of the transitions, the game has approximate equilibria.

**Proof:** Define a member of $\mathcal{R}$ to be **problematic** if the assumption of Theorem 1’ holds. We proceed exactly as the proof of Theorem 1, except that for all problematic $R$ we incorporate into the $S_\sharp$ transition $q^R_{R}$ all the $R$ exits not inside of $D_R$ or not created from a combination of an $a \in A_R$ with a move used in $y_C$. Recalling that $q^R_{R}$ is the difference between $q^d_{R}$ and $p^*_{R}$ by Lemma 3.7 we still have that $\nu^d_{\sharp}(q^d_{R})(v^2_{R}(q^d_{R}) - r^2_{\sharp}(s_R))$ is below $\hat{\epsilon}$. Due to Condition 2 and Lemma 3.7 we have the other requirement for applying Lemma 3.9. We assume that $\mathcal{T}$ is the subset of $\mathcal{R}$ that has been polarized.

Define a situation $s^w$ at a state $s$ to be **timed** if there is a natural number $m$ such that $s^w$ is determined by the present state $s$ and the previous situations and moves in the last $m$ stages. A normal situation is timed, but the converse doesn’t hold.

We keep the same situations $s^e$, $s^f$ and $s^g$ from the proof of Theorem 1. The stationary strategies for all the $s^g$ and all the $s^e$ other than a representative $s^e_R$ are defined in the same way, and in a non-problematic $R$ the stationary strategies for $s^f$ are also the same.

For every polarized $R \in \mathcal{T}$ and $t \in D_R$ we create a timed situation $t^h$. When a situation $s^e_R$ is reached the strategies $(x_C,y_C)$ are performed, but instead of moving to a $t^f$ or $t^g$ there is motion to the timed situation $t^h$.

For non-problematic polarized $R \in \mathcal{T}$ we choose any $t \in D_R$ such that there is a Player One move $a$ at $t$ not used in $x_C$ and when paired with $y_C$ results in zero probability of leaving the set $R$. Create a frequency $f_a > 0$ and a number $n_t$ such that $f_a \sum_{i=0}^{n_t-1} (1-f_a)^i = 1 - \lambda_R$, where $\lambda_R$ is that quantity determined by the polarization, and such that for any distinct $u, v \in D_R$ the probability of using the move $a$ before moving from $u$ to $v$ is at least $1 - \epsilon^*$. For all the situations $s^h$ for $s \neq t$ the players act according to $(x_C,y_C)$ and at $t$ Player Two according to $y_C$ and Player One according to $(1-f_a)x_C + f_a1_a$. If on the $n_t$th visit to the situation $t^h$ the move $a$ was not made, then the situation following $t^h$ is some $u^g$. Otherwise if the move $a$ was used on any visit to the situation $t^h$ then the next situation is either some $u^f$ if an exit wasn’t used or some $u^e$ if an exit was used.

For problematic $R \in \mathcal{R}$, let $\pi_R \in \Delta(A_R)$ be the probability distribution on the $A_R$ that is generated conditionally by $(x,y_C)$. Choose a natural number.
$n_R$ and a stationary strategy $x^*_C$ for Player One so that with a start at $s_R$ the distribution on the moves $A_R$ is $\pi_R$ and for every pair $u, v \in D_R$ the probability of using a move in $A_R$ before moving from $u$ to $v$ is no more than $\epsilon^*$ and the probability of using some member of $A_R$ at or before the $n_R$th visit to the state $s_R$ is $1 - \lambda_R$. For the situations $t^h$ with $t \in D_R$ the players act according to $(x^*_C, y_C)$. If on the $n_R$th visit to the situation $s_R^h$ the move $a$ was not made, then the situation following $t^h$ is some $u^g$. Otherwise if a move in $A_R$ was used on any visit to a situation $t^h$ then the next situation is either $u^f$ (if an exit was not used) or $u^c$ (if an exit was used). At a situation $u^f$ the strategies $(x^*_C, y_C)$ are also used.

As with the proof of Theorem 1 we must show that the expected payoffs to Player $i$ from every situation $s^e$ is within $3.1\tau$ of $r_{x,y}^i(s)$. Given the proof of Theorem 1 the only additional argument needed concerns the use of exits in a problematic $R$ before the timed situations have been reached. This did not present a problem in the proof of Theorem 1 because they were the same exits used in the situations $t^f$ and performed with the same distributions. If we can show that the total probability of their occurrence cannot exceed $\tau/10$, then we get our result by ignoring their influence. Indeed in the Markov chain defined on $S_\sharp$ the absorption rate of $s_R$ for a problematic $R$ is at least $\rho\hat{\mu}/(2Q_1)$. By Lemma 3.9 this absorption rate does not fall below $\rho\hat{\mu}/(2Q_1)(3|N|)|\mathcal{N}|$ after polarization. Since this quantity is still very large compared to $\epsilon^*$, the maximal probability of using such an exit before a timed situation is reached, we can indeed ignore these exits. (We leave the formal argument using Section 3 to the reader.)

The situations defined above are not normal and thus do not generate a stochastic game, preventing a direct application of Corollaries 4.3 and 4.4. Therefore we must perceive the situations $\{s^h | s \in R\}$ for $R \in T$ as subgames. Concerning the behavior of Player One, we view the entire process up until the $n_R$th visit to the state $s_R$ or the $n_t$th visit to $t$ as a single decision – whether or not to use a move in $A_R$ and if so then which one. This places Player One’s decisions back into the context of Corollary 4.4.

Concerning the behavior of Player Two, the matter is more complex. Player Two could have an influence on the payoffs by altering the strategy $y_C$. Strictly speaking the context would be no longer that of a harmonic function on a time homogeneous Markov chain – the expected payoff to Player Two at a state corresponding to a situation $t^h$ would be changing
over time. However Player Two’s ability to gain or lose in expected payoff is conditioned on the use of a move of Player One in $A_R$ — this is modeled by a time homogeneous Markov chain and therefore Proposition 4.2 is sufficient for the conclusion.

**Theorem 2'** The conditions of Theorem 1’ are satisfied always.

**Proof:** Let $(x, y) \in X \times Y$ and $(x_C, y_C)$ be a solution given by Parts 1, 2, and 3 of Theorem 2 for a polarized $R \in R$ and we assume that conditions of Theorem 1’ hold for $R$ (meaning that $R$ is problematic).

1) Consider the strategies played at any $t \in D_R$. Suppose for the sake of contradiction that there is a state $u \in R \setminus D_R$ where an importance of at least $\beta \tilde{e} \tilde{e}/3|R|$ occurs from exits at $u$. Consider the moves that were removed from $y^t$ to make $y^t_C$. By Lemma 5.5 at any $t \in D_R$ no more than $\frac{|R|}{3\tilde{e}} + \beta \frac{mL}{(L-1)Q_1}$ of the transition $P^t(u, v)$ was removed to make $y^t_C$ from $y^t$. On the other hand, given that every move of Player One removed from $x^t$ to make $x^t_C$ would have created an exit against some move in $y^t_C$, we must also conclude from the rare use of an exit that no more than $2\epsilon^* Q_1$ of the transition in $P^t_{x,y}(t, u)$ came from such a Player One move. From Lemma 3.3 we have that $u$ is in $D_R$, a contradiction.

2) Assume that $\nu_{x,y}^t \geq \beta \tilde{e} \tilde{e}/(3|N^t|m)$ for some principle move $a$ of Player One at $t \in D_R$. Suppose for the sake of contradiction that the probability of reaching any absorbing state from this principle move is altered by a factor of more than $\beta \tilde{e} \tilde{e}/2$ by the change from $y$ to $y_C$. This means that $\nu_{x,y}(a, b)$ is at least $\frac{\tilde{e}^2 \tilde{e}^2 \beta^2}{6|N^t|m^2}$ for at least one move $b$ that was removed to make $y^t_C$ from $y^t$. We must conclude from Lemma 5.5 that $\frac{\beta^2 \tilde{e}^2 \tilde{e}^2 \beta^2 a(t)}{6|N^t|m^2} \leq 2\epsilon^* Q_1 \frac{Q_1}{Q_1(L-1)}$, a contradiction to the definition of $Q_1$.

The final claim follows now from the argument in part 4 of the proof of Theorem 2, showing that $v^t_{x,y}(a)$ is very close to the value of $r^t$ for all primary moves.

In the proof of Theorem 1’ we could eliminate the argument that exits performed before reaching a timed situation in a problematic set are irrelevant if we had a more powerful Markov chain result (that combines the condition of Lemma 3.3 with the conclusion of Lemma 3.2) or we use Vieille’s approach to “communication sets” (Vieille 2000c), showing how one can move through a set $R$ with no danger of leaving it.

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8 Countably many states

On the technical side, the problem with applying either our or Vieille’s proof to countably many states lies in the finite state space assumption that given any stationary strategies for the players and any positive $\delta$ there will exist a $\delta$ perturbation of this strategy that is absorbing.

A strategy for finding a counter-example could be following. Construct an infinite sequence of games $\Gamma_0, \Gamma_1, \ldots$ that are positive recursive for both players corresponding to increasing finite sets $S_0 \subseteq S_1 \subseteq \ldots$ of non-absorbing states such that for every $i \geq 0$ and $j \geq i$ the moves and their induced motions inside of $S_i$ are the same for all games $\Gamma_j$. Construct a countable state space by having the game start at $s_0$, define the state space on the $i$th stage to be the space $S_i$, and declare that absorption occurs on stage $i$ if an absorbing state of the game $\Gamma_i$ is reached. Furthermore, give both players the ability to force the game to absorption in the new countable state space game. Desirable may be games $\Gamma_i$ such that with large $i$ the approximate equilibrium behavior of $\Gamma_i$ keeps the non-absorbing play most of the time close to the set $S_0$ and the minimal number of stages necessary to reach an absorbing state in the game $\Gamma_i$ starting from any $s_0 \in S_0$ goes to infinity as $i$ goes to infinity. Otherwise if we allow that absorbing states are reachable quickly from all non-absorbing states, to avoid convergence toward large sub-games of essentially equivalent states it may be desirable if reaching an absorbing state of $\Gamma_i$ on the $i$th stage of play does not mean certain absorption but rather a positive probability of absorption mixed with a positive probability of starting the game over at $s_0 \in \Gamma_0$.

There are many ways for a game to have a countable state space but be played essentially on finitely many situations, for example games that break down into sequences of sub-games played essentially on finite state spaces. Also to be avoided are structures that are formally countable in size but do not exploit the full potential of what it means to have infinitely many positional possibilities. We believe that the best candidates for a counter-example will incorporate the concept of a random walk on arbitrarily many positions, as presented in our introduction. However, to avoid operator approaches similar to that of the Maitra and Sudderth proof we believe that there must be a conflict by both players between exploiting their positions and controlling the behavior of the other player. For this and other reasons, we believe that the non-absorbing states must have a structure more complex.
than $\mathbb{Z}$, for example involving joint random walks on the two dimensional lattice $\mathbb{Z}^2$.

**Additional Acknowledgment:** The author thanks Cafe Europe on Azzahra Street in Jerusalem for the many hours he worked on the proof at the cafe.

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