COVID-19 vaccines reduce the risk of SARS-CoV-2 reinfection and hospitalization: Meta-analysis

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The addictive protection against SARS-CoV-2 reinfection conferred by vaccination, as compared to natural immunity alone, remains to be quantified. We thus carried out a meta-analysis to summarize the existing evidence on the association between SARS-CoV-2 vaccination and the risk of reinfection and disease. We searched MedLine, Scopus and preprint repositories up to July 31, 2022, to retrieve cohort or case-control studies comparing the risk of SARS-CoV-2 reinfection or severe/critical COVID-19 among vaccinated vs. unvaccinated subjects, recovered from a primary episode. Data were combined using a generic inverse-variance approach. Eighteen studies, enrolling 18,132,192 individuals, were included. As compared to the unvaccinated, vaccinated subjects showed a significantly lower likelihood of reinfection (summary Odds Ratio—OR: 0.47; 95% CI: 0.42–0.54). Notably, the results did not change up to 12 months of follow-up, by number of vaccine doses, in studies that adjusted for potential confounders, adopting different reinfection definitions, and with different predominant strains. Once reinfected, vaccinated subjects were also significantly less likely to develop a severe disease (OR: 0.45; 95% CI: 0.38–0.54). Although further studies on the long-term persistence of protection, under the challenge of the new circulating variants, are clearly needed, the present meta-analysis provides solid evidence of a stronger protection of hybrid vs. natural immunity, which may persist during Omicron waves and up to 12 months.

KEYWORDS
SARS-CoV-2, COVID-19, vaccination, meta-analysis, Omicron (B.1.1.529), reinfection
Introduction

Clarifying the frequency and predictors of SARS-CoV-2 reinfections is crucial to determine the course of the pandemic, and to optimize restriction and vaccination policies (1–3). Solid evidence is currently available on the frequency of reinfections after the emergence of the Omicron variant: a recent proportion meta-analysis including 15 million subjects recovered from a first infection estimated an overall reinfection rate of 3.3% in the first 3 months of Omicron predominance, likely increasing (2). However, the potential additive protection conferred by hybrid immunity, generated by the combination of prior infection and vaccination, as compared to the sole natural immunity, remains to be fully disclosed (4, 5). A few population-based studies suggested that reinfection is less likely in vaccinated vs. unvaccinated subjects, but the magnitude of the association varied across studies, which differed for patients’ characteristics, exposure risk, type of SARS-CoV-2 vaccine received, definition of reinfection adopted, and extent of measured confounding (4, 6–8). In a recent meta-analysis, the overall reinfection rate among vaccinated subjects was quantified to be as low as 0.32%, as compared to 0.74% among previously infected, unvaccinated individuals, but these estimates were obtained from raw, unadjusted data (2). Additionally, only limited data are available on the time course of natural and hybrid immunity (9), and the extent of its waning, particularly due to Omicron infections, is not yet well characterized (4, 9).

We carried out a meta-analysis to summarize the existing evidence from adjusted analyses on the association between SARS-CoV-2 vaccination and reinfection, in subjects who recovered from a first episode of SARS-CoV-2 infection.

Methods

Bibliographic search, data extraction and quality assessment

The reporting of this meta-analysis was guided by the standards of the Preferred Reporting Items for Systematic Review and Meta-Analysis (PRISMA) 2020 Statement (10). We searched MedLine and Scopus databases, up to July 31, 2022, for studies evaluating the risk of SARS-CoV-2 reinfection (either asymptomatic or symptomatic and requiring hospital admission) among vaccinated subjects of all ages (with hybrid immunity resulting from a combination of natural and vaccine immunization), vs. unvaccinated subjects (with natural immunity only). Vaccinated subjects were defined as those receiving ≥1 dose of the COVID-19 vaccines currently approved ≥14 days before the reinfection. The following search strategy was adopted, without language restrictions: (coronavirus∗ or coronavirus∗ or coronavirusae∗ or Coronavirus∗ or Coronavirus19 or Wuhan∗ or Hubei∗ or Huanan or “2019-nCoV” or 2019nCoV or nCoV2019 or “nCoV-2019” or “COVID-19” or COVID19 or “WN-CoV” or WNCov∗ or “HCoV-19” or HCoV19 or CoV or “2019 novelc∗” or Ncov∗ or “n-cov” or “SARS-CoV-2” or “SARSCoV-2” or “SARSCoV2∗” or SARSCoV2 or SARSCov19 or “SARS-CoV19” or “SARS-CoV-19” or “SARS-CoV” or “SARS-corona” or “Ncorona” or “Ncoronavirus” or NcovWuhan∗ or NcovHubei∗ or NcovChina∗ or NcovChinese∗) AND (reinfection∗ or reinfection∗ or second episode or recurrence∗ or recrudescence∗ or relapse∗ or ROCOVID19) (2). The reference lists of reviews and retrieved articles was also screened, for additional pertinent papers (11).

Inclusion criteria were: (a) cohort or case-control design; (b) laboratory confirmation of SARS-CoV-2 initial episode through a positive reverse-transcriptase polymerase chain reaction (RT-PCR) test, and/or an initial positive serology investigated with the use of an anti-trimeric spike IgG enzyme-linked immunosorbent assay (ELISA) (12); (c) data available to compare SARS-CoV-2 reinfection by vaccination status in subjects who recovered from a primary infection; (d) explicit reinfection definition criteria. In accordance with CDC (12), a reinfection was defined by the presence of:

(a) two positive PCR samples detected ≥45 days apart with ≥1 negative RT-PCR test collected between the first and second episode (13), and/or confirmation of infection with two different phylogenetic strains by viral genomic sequencing;

(b) two positive PCR samples detected ≥45 days apart in subjects with a symptomatic second episode or in close contact with a laboratory-confirmed COVID-19 case (12);

(c) a positive PCR test ≥45 days after the first positive serology (detection of anti-SI domain of spike protein IgG antibodies using an enzyme-linked immunosorbent assay—ELISA) (12, 14).

Each included article was independently evaluated by 2 reviewers (MEF, CAM), who extracted the main study characteristics and measures of effect. In case of discrepancies in data extraction, a third author was contacted (LM), and consensus achieved through discussion.

Individual study quality was evaluated using an adapted version of the Newcastle Ottawa Quality Assessment Scale, assessing the comparability across groups for confounding factors, the appropriateness of outcome assessment, length of follow-up and missing data handling and reporting (15).

Data analysis

The units of the meta-analysis were single comparisons of vaccinated vs. unvaccinated subjects in predicting (a)
### Table 1: Characteristics of the included studies.

| No. | References | Journal | Country | Design | Population | % vacc. | Mean age (SD) | Mean f-up (days) | Dominant strain | Reinfecion definition and time-lag | Raw data | Covariates |
|-----|------------|---------|---------|--------|------------|---------|---------------|----------------|----------------|-----------------------------------|----------|------------|
| 1   | Bager et al.\(^{1, b}\) (27) | Lancet Infect Dis | Denmark | Cohort | General | 65.8 | 31.0 (27.4) | 120 | Delta | 2 PCR + > 60 days | 783/80,426 vs. 1103/69,885 | Raw data extracted |
| 2   | Bager et al.\(^{a2, b}\) (27) | Lancet Infect Dis | Denmark | Cohort | General | 81.2 | 29.0 (18.5) | 120 | Omicron | 2 PCR + > 60 days | 1520/31,403 vs. 622/7266 | Raw data extracted |
| 3   | Cavanaugh et al. (21) | MMWR | USA | Case-control | General | 20.3 | NR | NR | NR | PCR + /Ag test May–Jun 21 (1st episode: Mar–ec 20) | 67/275 vs. 179/463 | Age, gender, time from 1st infection |
| 4   | Cerqueira-Silva et al. (26) | Lancet Infect Dis | USA | Case-control | General | 35.5 | 36.0 (11.1) | 60 | Gamma | 2 PCR + > 90 days | 6584/59,064 vs. 14,566/97,856 | Comorb, time from 1st infection, severity of 1st infection |
| 5   | Eythorsson et al. (6) | JAMA Netw Open | Iceland | Cohort | General | 25.5 | 34.0 (19.0) | 287 | Omicron | 2 PCR + > 60 days | 320/2,938 vs. 1,007/8,598 | Age, gender, time from 1st infection |
| 6   | Flacco et al. (28) | Front Public Health | Italy | Cohort | General | 43.5 | 41.6 (21.9) | 277 | Omicron | 2 PCR + ≥ 45 days (≥ 1 PCR−) | 386/88,576 vs. 343/30,690 | Age, gender, comorb, severity of 1st infection |
| 7   | Hall et al. (29) | Lancet | UK | Cohort | HCW | 47.5 | 45.6 (14.2) | 275 | NR | 2 PCR + ≥ 90 days + serology/genomic | NR | Age, gender, ethnicity, time from 1st infection, workplace, contact frequency |
| 8   | Hammerman et al. (7) | New Engl J Med | Israel | Cohort | General | 56.0 | 39.3 (17.1) | 270 | Delta | 2 PCR + > 90 days | 354/83,356 vs. 2,668/65,676 | Age, gender, comorb. ethnicity, socio-economic status |
| 9   | Jang et al. (30) | J Med Virol | Korea | Cohort | General | 76.1 | NR | 242 | Omicron | 2 PCR + ≥ 45 days | 19,943/12,270,241 vs. 19,513/3,638,932 | Age, gender, strain immunologic status |
| 10  | Levin-Rector et al. (22) | Clin Infect Dis | USA | Case-control | General | 54.4 | NR | NR | Delta | 2 PCR + > 90 days | 965/5,228 vs. 1,436/4,376 | Age, gender, time from 1st infection |
| 11  | Lewis et al.\(^{1, c}\) (31) | JAMA Netw Open | USA | Cohort | General | 51.2 | 35.0 (20.7) | 225 | Delta | 2 PCR + > 90 days | 298/52,683 vs. 1,105/41,833 | Age, gender, time from and severity of 1st infection |
| 12  | Lewis et al.\(^{2, c}\) (31) | JAMA Netw Open | USA | Cohort | HCW | 66.3 | 41.0 (17.0) | 225 | Delta | 2 PCR + > 90 days | 47/2,131 vs. 227/746 | Age, gender, time from and severity of 1st infection |
| 13  | Malhotra et al. (32) | JAMA Netw Open | India | Cohort | HCW | 75.3 | 36.6 (10.3) | 233 | Delta | 2 PCR + ≥ 90 days | 56/1,445 vs. 60/472 | Age, gender, work category |
| 14  | Medic et al. (4) | Lancet Reg Health | Serbia | Case-control | General | 46.2 | 45.9 (18.7) | 340 | Omicron | Rapid Ag test or 2 PCR + ≥ 90 days | 3,404/10,220 vs. 3,815/11,417 | Age, gender, comorb., time from 1st infection |
| 15  | Murugesan et al. (33) | PloS One | India | Cohort | HCW | 76.9 | 33.7 (10.9) | 259 | Delta | 2 PCR + ≥ 90 days | 12,791 vs. 16,658 | Raw data extracted |
| 16  | Nisha et al. (34) | J Fam Commun Med | India | Cohort | HCW | 36.3 | 30.3 (10.5) | 270 | NR | 2 PCR + > 90 days (≥ 1 PCR−) | 103/1,684 vs. 24/225 | Age, gender, comorb, work category |

(Continued)
### TABLE 1 (Continued)

| No. | References | Journal | Country | Design | Population | % vacc. Mean age (SD) | Mean f-up (days) | Dominant strain | Reinfection definition and time-lag | Covariates |
|-----|-------------|---------|---------|--------|------------|-----------------------|----------------|-----------------|-------------------------------------|------------|
| 17  | Nordstrom et al. (8) | Lancet Infect Dis | Sweden | Cohort | General | 50.0 38.8 (17.9) | 60 | Delta PCR + | Dec 20-Oct 21 (1st episode before 24 May 21) | Age, gender, comorb., time from 1st infection, marital status, work category |
| 18  | Nunes et al. (23) | Vaccines | South Africa | Case-control | HCW | 80.0 37.4 (9.2) | NR | Omicron 2 PCR + | > 90 days | Study site |
| 19  | Plumb et al. (24) | MMWR USA | USA | Case-control | General | 48.4 NR | NR | Omicron 2 PCR + | > 90 days | Age, gender, race, time from 1st infection |
| 20  | Plumb et al. (24) | MMWR USA | USA | Case-control | General | 48.4 NR | NR | Omicron 3 PCR + | > 90 days | Age, gender, race, time from 1st infection |
| 21  | Spicer et al. (25) | J Pediatric USA | USA | Case-control | General | 20.5 15.1 (1.7) | 246 | Delta PCR + | 20/855 vs. 342/3,307 | Raw data extracted |

SARS-CoV-2 reinfection; (b) severe COVID-19 disease—requiring hospital admission with no use of an intensive care unit; (c) critical/lethal COVID-19 disease—requiring admission in an intensive care unit and/or causing death (2). The likelihood of each outcome was assessed: (a) using ≥ 45 days as the minimum time-lag between two positive episodes; (b) adopting a more stringent time-lag of 90 days (2); (c) including only studies with adjusted estimates. When data were available, we also performed several additional meta-analyses stratified by: (d) number of vaccine doses (“fully vaccinated” subjects—those receiving ≥ 2 doses of mRNA-1273, BNT162b2, ChAdOx1 nCoV-19, BBV152, BBIBP-CorV, Gam-COVID-Vac, CoronaVac, or 1 dose of JNJ-78436735 ≥ 14 days before reinfection—or “partially vaccinated” subjects—those receiving 1 dose of mRNA-1273, BNT162b2, ChAdOx1 nCoV-19, BBV152, BBIBP-CorV, Gam-COVID-Vac, or CoronaVac ≥ 14 days before reinfection—vs. unvaccinated) (13). When data were available, we also extracted separate estimates for those who received 3 doses of mRNA-1273, BNT162b2, ChAdOx1 nCoV-19, BBV152, BBIBP-CorV, Gam-COVID-Vac, or CoronaVac vaccines (“boosted subjects”); (e) time between first episode and reinfection (<6 vs. ≥ 6 months); (f) dominant viral strain (Delta or Omicron); (g) exposure risk (healthcare workers or general population); (h) study design (cohort or case-control).

Data were combined using a random-effect generic inverse variance approach (16, 17), in order to account for between-study heterogeneity (18). If a study reported the results of different multivariable models, the most stringently controlled estimates (those from the model adjusting for more factors) were extracted. If different models controlled for the same number of covariates, the model containing the most clinically meaningful covariates was used for the analysis (19). When a study only reported separate estimates by vaccine dose, the overall estimate of risk was computed from the separate relative risks using the fixed-effect model for generic inverse-variance outcomes (19).

Between-study heterogeneity was quantified using the I² statistic. Potential publication bias was assessed graphically, using funnel plots [displaying the Odds Ratios—ORs from individual comparisons vs. their precision (1/SE)], and formally, using Egger’s regression asymmetry test (16).

All meta-analyses were performed using RevMan software, version 5.3 [The Cochrane Collaboration, (20)].

### Results

Of the 3,470 papers initially retrieved, seven case-control (4, 21–26) and 11 cohort studies (6–8, 27–34) were included in the analyses [Supplementary Figure 1 and Supplementary Table 1]. Three studies contributed with two dataset (24, 27, 31), as the same publication provided separate data for healthcare workers and the general population (31), and for Delta and Omicron...
waves (24, 27): this led to a total of 21 datasets that were included in the analyses (Table 1).

Six studies were carried out in Europe (4, 6, 8, 24, 27–29), six in the USA (21, 22, 24–26, 31), five in Asia (7, 30, 32–34) and one in South Africa (23). Thirteen studies evaluated the general population (4, 6–8, 21, 22, 24–28, 30, 31), and six assessed the healthcare workers (23, 29, 31–34). In most studies, the analyses were adjusted for age, gender, and comorbidities, as a minimum set of potential confounders of the association between vaccination status and reinfections (4, 6–8, 21, 23, 26, 28–32, 34).

The mean age of the participants ranged from 15 to 46 years, and the mean follow-up ranged from a minimum of 60 up to 340 days. In 13 studies (4, 7, 21–26, 29, 31–34) the minimum time-lag between infection and reinfection was set at 90 days, and only three (28, 29, 34) strictly followed the CDC criteria to define a reinfection (≥1 intermediate negative PCR and/or viral genomic sequencing) (12). Most reinfections were reported during the Delta (7, 8, 22, 24, 25, 27, 31–33) and the Omicron waves (4, 6, 23, 24, 27, 28, 30).

The methodological characteristics of the included studies are summarized in Table 2: the selection of the cohort of patients, the ascertainment of the exposure, and the evaluation of the comparability of subjects were adequate in all studies, while 15 out of 18 adequately addressed the items pertaining to outcome assessment and follow-up (length and missing data).

### Table 2 Methodological quality of the included studies according to the Newcastle Ottawa Scale.

| References                   | Selection (Max. score 4) | Comparability (Max. score 2) | Outcome (Max. score 3) |
|------------------------------|--------------------------|------------------------------|------------------------|
| Bager et al. [27]            | 4                        | 2                            | 3                      |
| Cavanaugh et al. [21]        | 4                        | 2                            | 3                      |
| Cerqueira-Silva et al. [26]  | 4                        | 2                            | 3                      |
| Eythorsson et al. [6]        | 4                        | 2                            | 3                      |
| Flacco et al. [28]           | 4                        | 2                            | 3                      |
| Hall et al. [29]             | 4                        | 2                            | 3                      |
| Hammerman et al. [7]         | 4                        | 2                            | 3                      |
| Jang et al. [30]             | 4                        | 2                            | 3                      |
| Levin-Rector et al. [22]     | 4                        | 2                            | 2                      |
| Lewis et al. [31]            | 4                        | 2                            | 3                      |
| Malhotra et al. [32]         | 3                        | 2                            | 3                      |
| Medic et al. [4]             | 4                        | 2                            | 3                      |
| Murugesan et al. [33]        | 4                        | 2                            | 3                      |
| Nisha et al. [34]            | 4                        | 2                            | 3                      |
| Nordstrom et al. [8]         | 4                        | 2                            | 2                      |
| Nunes et al. [23]            | 3                        | 2                            | 2                      |
| Plumb et al. [24]            | 4                        | 2                            | 3                      |
| Spicer et al. [25]           | 4                        | 2                            | 2                      |

Twenty-one datasets including a total of 18,132,192 individuals were included in the overall meta-analysis comparing the risk of SARS-CoV-2 reinfection in vaccinated vs. unvaccinated subjects (Table 3) (4, 6–8, 21–34). In 20 out of 21 datasets, the vaccinated subjects were significantly less likely to be reinfected, with a summary OR of 0.47 (95% confidence interval—CI – 0.42–0.54) (Figure 1). When the only study reporting a significantly higher risk among vaccinated subjects (and no data on underlying comorbidities) was excluded (6), the estimates were virtually identical (OR: 0.45; 95% CI: 0.39–0.50).

Also, the results did not substantially change after the exclusion of the three studies with unadjusted estimates (OR: 0.47; 95% CI: 0.39–0.56) (25, 27, 33), and when only the 17 datasets with a more conservative time-lag of 90 days were considered (OR: 0.47; 95% CI: 0.41–0.54) (4, 7, 21, 23, 24, 26, 29, 31–34).

When the analyses were stratified by number of doses, the summary OR of reinfection was lower among fully vaccinated than partially vaccinated subjects (summary OR 0.45 and 0.58, respectively). The confidence intervals, however, largely overlapped. In the analyses restricted to the subjects who received three doses (a booster dose), the summary OR was comparable to that of the fully vaccinated individuals (OR: 0.46; 95% CI: 0.29–0.73). As shown in Table 3, the association between vaccination and reinfection did not show a substantial variation by length of follow-up: the summary OR of the studies with a follow-up shorter than 6 months (OR: 0.52; 95% CI: 0.40–0.67) was comparable with the OR (0.45; 95% CI: 0.34–0.59) of the studies with a longer follow-up (up to 340 days).

The likelihood of a reinfection remained significantly lower among vaccinated subjects both in the studies that were carried out during Delta predominance (summary OR: 0.40; 95% CI: 0.31–0.50) (7, 8, 19, 22–24, 27–29) and during Omicron predominance (OR: 0.58; 95% CI: 0.48–0.60) (2, 4, 6, 23, 24, 27, 30). Again, in the analyses stratified by risk of exposure (general population or healthcare workers) and by study design (cohort or case-control) the likelihood of reinfection was comparably, significantly lower among vaccinated subjects, with summary ORs ranging from 0.44 to 0.54, and overlapping confidence intervals.

The Egger test was not significant (p = 0.3), and the funnel plot displaying the ORs of the individual comparisons vs. the logarithm of their SE (precision) did not show asymmetry, suggesting the absence of publication bias (Supplementary Figure 2).

A total of seven datasets and 2,312,703 individuals provided specific data and were included in the meta-analysis comparing the risk of severe/lethal COVID-19 of the vaccinated vs. the unvaccinated subjects (8, 22, 24, 26, 29, 32). Compared with the unvaccinated, those receiving ≥ 1 dose were significantly less likely to develop a severe disease, once reinfected (OR: 0.45; 95% CI: 0.38–0.54—Table 3 and Figure 2). The risk remained comparably and
TABLE 3 Risk of SARS-CoV-2 reinfection and severe/critical COVID-19 among vaccinated vs. unvaccinated subjects, overall, and stratified by definition of reinfection, number of vaccine doses, length of follow-up, predominant strain, study design and risk exposure.

| Analyses                                                                 | Pooled estimates          | Raw data\(^b\)          |
|--------------------------------------------------------------------------|----------------------------|--------------------------|
|                                                                          | N. datasets (total sample)
|                                                                          | OR (95% CI)   | P-value | I², % | No. of events Vacinated | No. of events Unvaccinated |
| SARS-CoV-2 reinfection—all studies (4, 6, 8, 21–34)                      | 21 (18, 132, 192)         | 0.47 (0.42 – 0.54)       | < 0.001 | 98     | 37,440                  | 13,462,121                |
| - Adjusted estimates only (4, 6, 8, 21–24, 26, 28–32, 34)               | 17 (17, 937, 601)         | 0.47 (0.41 – 0.54)       | < 0.001 | 98     | 35,105                  | 13,348,646                |
| 1. Time-lag ≥ 90 days\(^c\)                                             | 15 (373, 109)             | 0.44 (0.36 – 0.54)       | < 0.001 | 97     | 13,411                  | 223,473                  |
| - Adjusted estimates only (4, 21–24, 26, 29, 31, 32, 34)                | 13 (367, 498)             | 0.46 (0.37 – 0.56)       | < 0.001 | 97     | 13,379                  | 221,827                  |
| 2. Number of vaccine doses:\(^d\)                                       |                            |                          |         |        |                         |                          |
| - Partially vaccinated subjects (4, 8, 23, 24, 26, 28–32)                | 11 (5, 248, 720)          | 0.58 (0.44 – 0.77)       | 0.004   | 98     | 5,820                   | 729,103                  |
| - Fully vaccinated subjects (4, 8, 21–24, 26, 28–32)                     | 13 (7, 036, 021)          | 0.45 (0.40 – 0.50)       | < 0.001 | 95     | 28,508                  | 12,521,565               |
| - Boosted subjects (3 doses) (4, 24, 30)                                | 4 (11, 365, 430)          | 0.46 (0.29 – 0.73)       | 0.001   | 99     | 1,675                   | 7,709,207                |
| 3. Length of follow-up:                                                  |                            |                          |         |        |                         |                          |
| - <6 months (< 120 days)—all studies (8, 26, 27)                        | 4 (1, 876, 028)           | 0.52 (0.40 – 0.67)       | < 0.001 | 99     | 9,964                   | 935,957                  |
| - Adjusted estimates only (8, 26)                                       | 2 (1, 603, 758)           | 0.47 (0.30 – 0.74)       | < 0.001 | 99     | 7,661                   | 824,128                  |
| - ≥6 months (≥ 120 days)—all studies (4, 6, 7, 25, 28–34)               | 12 (16, 317, 474)         | 0.45 (0.34 – 0.59)       | < 0.001 | 98     | 24,943                  | 12,514,920               |
| - Studies with adjusted estimates only (4, 6, 7, 28–32, 34)             | 10 (16, 311, 863)         | 0.47 (0.35 – 0.63)       | 0.05    | 99     | 24,911                  | 12,515,274               |
| 4. Predominant viral strain:                                             |                            |                          |         |        |                         |                          |
| - Delta variant (B.1.617.2)—all studies (8, 22, 24, 25, 27, 31–33)      | 10 (1, 948, 597)          | 0.40 (0.31 – 0.50)       | < 0.001 | 97     | 4,099                   | 994,162                  |
| - Adjusted estimates only (8, 22, 24, 31, 32)                           | 7 (1, 792, 675)           | 0.38 (0.30 – 0.49)       | < 0.001 | 96     | 3,284                   | 912,090                  |
| - Omicron variant (B.1.529)—all studies (4, 6, 23, 24, 27, 28, 30)     | 7 (16, 107, 318)          | 0.58 (0.48 – 0.70)       | < 0.001 | 97     | 26,587                  | 12,406,936               |
| - Adjusted estimates only (4, 6, 23, 24, 28, 30)                        | 6 (15, 951, 396)          | 0.59 (0.48 – 0.73)       | < 0.001 | 96     | 25,772                  | 12,324,864               |
| 5. Risk of exposure:                                                     |                            |                          |         |        |                         |                          |
| - General population—all studies (4, 8, 21, 22, 24–28, 30, 31)          | 15 (18, 123, 901)         | 0.47 (0.41 – 0.53)       | < 0.001 | 98     | 37,179                  | 13,455,954               |
| - Adjusted estimates only (4, 8, 21, 22, 24, 26, 28, 30, 31)            | 11 (17, 930, 759)         | 0.46 (0.37 – 0.55)       | < 0.001 | 98     | 34,856                  | 13,343,270               |
| - Healthcare workers—all studies (23, 29, 31–34)                         | 6 (8, 291)                | 0.50 (0.41 – 0.61)       | < 0.001 | 0      | 261                     | 6,167                    |
| - Adjusted estimates only (23, 29, 31, 32, 34)                          | 5 (6, 842)                | 0.49 (0.40 – 0.61)       | < 0.001 | 0      | 249                     | 5,376                    |

(Continued)
### TABLE 3 (Continued)

| Analyses | N. datasets (total sample) | OR (95% CI) | P-value | $I^2$, % | No. of events | Vaccinated subjects | No. of events | Unvaccinated subjects |
|----------|---------------------------|-------------|---------|---------|--------------|-------------------|--------------|----------------------|
| **Pooled estimates** | | | | | | | | |
| 6. Study design: | | | | | | | | |
| - Cohort—all studies (6, 8, 25, 27–34) | 14 (18, 014, 945) | 0.44 (0.36 – 0.54) | < 0.001 | 98 | 24,919 | 13,381,593 | 29,000 | 4,633,352 |
| - Adjusted estimates only (6, 8, 28–32, 34) | 7 (117, 247) | 0.54 (0.48 – 0.61) | < 0.001 | 89 | 12,521 | 80,528 | 105,598 | 36,719 |
| - Case-control—all studies (4, 21–24, 26) | 7 (2, 312, 703) | 0.45 (0.38 – 0.54) | < 0.001 | 91 | 1,411 | 1,536,917 | 2,657 | 775,786 |
| **Severe or critical/lethal COVID-19** | | | | | | | | |
| **1. Number of vaccine doses** | | | | | | | | |
| - Partially vaccinated subjects (5, 24, 26, 32) | 5 (982, 721) | 0.35 (0.21 – 0.60) | 0.02 | 91 | 474 | 48,447 | 3,693 | 498,250 |
| - Fully vaccinated subjects (5, 22, 24, 26, 32) | 6 (597, 193) | 0.34 (0.24 – 0.49) | < 0.001 | 93 | 1,629 | 296,197 | 3,620 | 300,996 |

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*a Three studies (34, 27, 31) contributed with more than one dataset, thus the number of references does not always match the number of datasets included in each analysis (see "Results" for further details).

*b Number of events/Total number of previously infected and vaccinated subjects vs. Number of events/Total number of previously infected and unvaccinated subjects.

*c The risk of SARS-CoV-2 reinfection was computed: (1) using $\geq 45$ days as the minimum time-lag between two positive episodes; (2) adopting a more stringent time-lag of 90 days (see Methods for further details).

*d Full vaccination was defined as complete vaccination with two doses of the same vaccine, and partial vaccination as one dose of vaccine (and sometimes booster doses).

*e Partially vaccinated subjects: 1 dose of mRNA-1273, BNT162b2, ChAdOx1 nCoV-19, BIV152, BIBP219, Gam-COVID-Vac, or CoronaVac and/or 14 days before the second dose. Fully vaccinated subjects: 2 doses of mRNA-1273, BNT162b2, ChAdOx1 nCoV-19, BIV152, BIBP219, Gam-COVID-Vac, or CoronaVac, and/or 14 days before the second dose.

**Severe COVID-19**: disease requiring hospital admission with no use of an intensive care unit; critical/lethal COVID-19: disease requiring admission in an intensive care unit and/or causing death. OR, Odds ratio; CI, confidence interval.
longer follow-up, the odds of reinfection were approximately 50% lower among the vaccinated. Inevitably, this information remains preliminary, as it is based upon studies in which the follow-up lasted up to 12 months, and the use of viral genomic sequencing was uneven.

These findings may offer a contribution to help planning tailored immunization strategies for previously infected subjects: if, on one side, the marked increase in the absolute number of reinfections with time is concerning, the significantly lower relative risk still observed among vaccinated subjects may be reassuring, thus vaccinating also this population may definitely play a role to control the pandemic (4). In this scenario, the strong protective effect exerted by a single dose (if confirmed during longer follow-up and toward different strains) might be taken into account when designing tailored vaccination schedules directed to lower-priority groups (4, 5). It should be also considered, however, that the degree of additional protection specifically conferred by further boosters (three or more doses) still remains uncertain, as our stratified meta-analyses did not show a clear benefit of a 3- vs. a 2-dose schedule.

The second main finding of the present meta-analysis was the significant reduction of the risk of hospitalization due to severe COVID-19 that was observed among the vaccinated subjects, either receiving one or more doses. This was crucial, as the primary aim of COVID-19 vaccination is to reduce the pressure on the healthcare systems preventing severe disease and hospitalization (37). Unfortunately, however, most of the studies included in the meta-analyses of this outcome were carried out before the emergence of Omicron strain. Therefore, this finding requires confirmation from more recent data with longer follow-up, as the large increase in the number of reinfections during the Omicron wave, and in turn the consequences on the healthcare systems still needs to be carefully evaluated.
In the first phases of the pandemic, there was uncertainty on the criteria to define a reinfection, especially on the time interval between the first and second episodes, and most initial studies defined a reinfection as a new PCR test occurring ≥ 90 days after complete resolution of the first infection (4, 7, 21–26, 29, 31–34). However, the CDC later expanded the definition, including also the subjects with COVID-19-like symptoms and detection of SARS-CoV-2 RNA ≥ 45 days since first infection (12). In the present analysis, we did not find substantial differences when a 90-day or a 45-day cutoff was adopted, suggesting that a low proportion of reinfections was missed using the longer threshold. Indeed, a recent cohort study reported a mean time between the first and second infection of 349 days, with less than 15% of the reinfections occurring in the first 6 months since the first episode (28).

Some limitations must be considered when interpreting the present findings. First, most meta-analyses showed an intermediate-to-high level of heterogeneity. However, a certain degree of heterogeneity across studies was inevitable, given the large variation in terms of setting and baseline patients characteristics. Also, when the analyses were repeated adopting a fixed approach, none of the results substantially differed (except for CIs, which were typically tighter). Second, although most studies provided analyses at least adjusted for age, gender, and several underlying comorbidities, some extent of residual confounding cannot be completely ruled out, as for any observational study (38). Third, the risk of reinfection could have been overestimated in several of the included studies adopting less stringent criteria to define a reinfection (2). Conversely, if previously infected people tended to seek fewer testing due to their presumed acquired natural immunity, the reinfection rate could have been underestimated (4). A sensitivity analysis based upon the average number of PCR tests as a proxy of health-seeking behavior would have increased the precision of our estimates (2), but these data were unfortunately not available. Fourth, it might have been interesting to evaluate if the results differed according to the sequence of events, whether vaccination was administered before or after the first infection. Unfortunately, however, the exact timeline of events could be determined only in two studies (4, 31), in which all the infections occurred before the start of the vaccination campaign.

Acknowledging these caveats, this meta-analysis showed that, among the subjects that recovered from a first SARS-CoV-2 infection, vaccination was associated with a significant and substantial reduction of the risk of both reinfection and severe COVID-19. This finding was confirmed when the analyses were adjusted for potential confounders, up to 12 months of follow-up, and after any vaccine dose. Further studies on the long-term persistence of protection, and assessing the reinfection and hospitalization rates under the challenge of the new circulating variants, are strongly warranted.

Data availability statement

The data presented in this study are available upon reasonable request from the corresponding author.

Author contributions

MF and LM: concept and design and statistical analysis. MF, CA, VB, and ER: acquisition, analysis, or interpretation of data. MF, CA, and LM: drafting of the manuscript. CD, PV, and LM: critical revision of the manuscript for important intellectual content. PV and LM: supervision. LM: full access to all the data in the study and take responsibility for the integrity of the data and the accuracy of the data analysis. All authors contributed to the article and approved the submitted version.
Conflict of interest

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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Supplementary material

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fmed.2022.1023507/full#supplementary-material

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