Blood glucose concentration and risk of pancreatic cancer: systematic review and dose-response meta-analysis

Wei-Chih Liao assistant professor, Yu-Kang Tu associate professor, Ming-Shiang Wu professor, Jaw-Town Lin professor, Hsiu-Po Wang professor, Kuo-Liong Chien professor

1Department of Internal Medicine, National Taiwan University Hospital, National Taiwan University College of Medicine, 7 Chang Shan South Road, Taipei 100, Taiwan; 2Institute of Epidemiology and Preventive Medicine, National Taiwan University, 17 Hsu Chow Road, Taipei 100, Taiwan; 3School of Medicine, Fu Jen Catholic University, 510 Zhongzheng Road, New Taipei City 242, Taiwan

Abstract

Objective To evaluate potential linear and non-linear dose-response relations between blood glucose and risk of pancreatic cancer.

Design Systematic review and dose-response meta-analysis of prospective observational studies.

Data sources Search of PubMed, Scopus, and related reviews before 30 November 2013 without language restriction.

Eligibility criteria Prospective studies evaluating the association between blood glucose concentration and pancreatic cancer. Retrospective and cross sectional studies excluded to avoid reverse causality.

Data extraction and synthesis Two reviewers independently extracted relevant information and assessed study quality with the Newcastle-Ottawa scale. Random effects dose-response meta-analysis was conducted to assess potential linear and non-linear dose-response relations.

Results Nine studies were included for analysis, with a total of 2408 patients with pancreatic cancer. There was a strong linear dose-response association between fasting blood glucose concentration and the rate of pancreatic cancer across the range of prediabetes and diabetes. No non-linear association was detected. The pooled rate ratio of pancreatic cancer per 0.56 mmol/L (10 mg/dL) increase in fasting blood glucose was 1.14 (95% confidence interval 1.06 to 1.22; P<0.001) without significant heterogeneity. Sensitivity analysis excluding blood glucose categories in the range of diabetes showed similar results (pooled rate ratio per 0.56 mmol/L increase in fasting blood glucose was 1.15, 95% confidence interval 1.05 to 1.27; P=0.003), strengthening the association between prediabetes and pancreatic cancer.

Conclusions Every 0.56 mmol/L increase in fasting blood glucose is associated with a 14% increase in the rate of pancreatic cancer. As prediabetes can be improved or even reversed through lifestyle changes, early detection of prediabetes coupled with lifestyle changes could represent a viable strategy to curb the increasing incidence of pancreatic cancer.

Introduction

Pancreatic adenocarcinoma is the most lethal cancer, with a five year survival rate of less than 5%.¹ The incidence and mortality rates of pancreatic cancer are increasing; it is the fifth and fourth leading cause of cancer deaths in the United Kingdom and the United States, respectively, and globally causes an estimated 227 000 deaths a year.² As about 85% of the tumors are already unresectable at diagnosis,³ prevention through modification of its risk factors is especially important.⁴ ⁵ Epidemiological evidence supports type 2 diabetes as a risk factor for pancreatic cancer,⁴ ⁶ and chronic hyperinsulinaemia and hyperglycaemia associated with type 2 diabetes have been proposed as the underlying mechanism. Experimental evidence suggests that insulin promotes proliferation and reduces apoptosis in pancreatic cancer cells, both directly and indirectly through increased bioavailability of insulin-like growth factor 1.⁴ ⁷ ⁸ Hyperglycaemia can also enhance proliferation⁹ ¹⁰ and invasion ability¹¹ of pancreatic cancer cells.¹² Hyperinsulinaemia and hyperglycaemia, however, are already present at the stage of prediabetes (blood glucose between normal and diabetes—that is, fasting blood glucose 5.6-6.9 mmol/L, post-load blood glucose 7.8-11.0 mmol/L, or haemoglobin A₁c 5.7-6.4%), which precedes type 2 diabetes.¹³ ¹⁴ Taken together, these observations...
suggest that prediabetes could also increase the risk of pancreatic cancer.

Whether prediabetes is a risk factor for pancreatic cancer has important implications for prevention. Prediabetes affects about 7.8% (344 million) of the world’s adult population, but changes in lifestyle (weight loss, dietary modification, and physical activity) could improve and even reverse this. Previous epidemiological studies examining the association between blood glucose and pancreatic cancer, however, yielded inconsistent results, and the knowledge gap remains wide. The association between prediabetes and pancreatic cancer was significant in one study but non-significant in four others; three studies had mixed results. No study assessed whether there is a threshold concentration of blood glucose that raises risk. Furthermore, those studies might have only modest statistical power as their categorisation of the continuous exposure (that is, glucose concentration) inevitably causes loss of information, and the small number of patients with pancreatic cancer in many glucose categories could further reduce their power. Collectively, whether prediabetes indeed increases the risk of pancreatic cancer remains controversial, and the dose-response relation between blood glucose concentration and risk has not been investigated.

Dose-response meta-analysis is a potential solution to the above problems as it enables evaluation of both linear and non-linear dose-response relations and pools multiple studies to offer greater statistical power. We therefore carried out a random effects dose-response meta-analysis to examine the relation between blood glucose concentrations and risk of pancreatic cancer.

Methods

Search strategy and selection criteria

This systematic review followed the guidelines in the MOOSE (Meta-analysis of Observational Studies in Epidemiology) statement. Two investigators (WCL and YKT) independently searched PubMed and Scopus and reviewed titles/abstracts in duplicate for studies that examined the association between blood glucose concentrations and pancreatic cancer from database inception to the end of November 2013 without language or date restrictions (see appendix 1 for the detailed study protocol with search strategies). We also manually searched bibliographies of included studies and related reviews for additional references. Studies that prospectively evaluated the association between blood glucose and pancreatic cancer were included for meta-analysis. As pancreatic cancer induces diabetes in about 40% of patients, we excluded retrospective and cross sectional studies to avoid reverse causality.

Data extraction and assessment for study quality

Two investigators (WCL and YKT) independently reviewed full manuscripts of eligible studies and extracted information into an electronic database, including author, publication year, country where the study was conducted, study design, sample size, duration of follow-up, methods of measurement/categorisation of blood glucose concentrations and outcome ascertainment, number of cases, rate ratios and 95% confidence intervals, and adjusted covariates. The same reviewers assessed study quality independently with the Newcastle-Ottawa scale. Disagreement was resolved by joint review of the manuscript to reach consensus.

Data synthesis and analysis

The outcome we analysed was the rate ratio of the incidence or mortality of pancreatic cancer; for pancreatic cancer incidence almost equals mortality. Random effects models were used for all meta-analyses to account for heterogeneity among studies. If only separate rate ratios for men and women were available in the original report, we pooled sex specific rate ratios using fixed effect models for subsequent meta-analysis when feasible. We used fasting blood glucose concentration as the exposure because it is the most reliable and convenient test for diagnosis of type 2 diabetes and was reported by most of the included studies. For studies that used haemoglobin A1c or post-load blood glucose concentration after an oral glucose tolerance test as the exposure, we converted these data by using the following method: the cut off fasting blood glucose for prediabetes (5.6 mmol/L) and diabetes (7.0 mmol/L) were assumed to be equivalent to the cut offs of haemoglobin A1c (5.7% and 6.5%) and post-load blood glucose (7.8 mmol/L and 11.1 mmol/L), respectively. This approach has been shown to have acceptable accuracy. For each of the included studies, we assigned the reported median or mean blood glucose concentration of each category as the category blood glucose concentration. When a study reported only the range of blood glucose for a category, we used the average value of the lower and upper bounds of that category. When the highest category was open ended, its category blood glucose concentration was calculated as the lower bound plus 1.5 times the width of the neighboring category. When the lowest category was open ended, its category blood glucose concentration was calculated as the average of the upper bound and 3.9 because the lower limit of fasting blood glucose is normally around 3.9 mmol/L.

We first summarised the rate ratios for the highest versus the lowest category of fasting blood glucose in included studies using the random effects meta-analysis proposed by DerSimonian and Laird (high v low meta-analysis). Potential small study bias was evaluated by funnel plots and by Egger’s test and Begg’s test. Heterogeneity was evaluated by I² and Cochran’s Q. For dose response meta-analysis, we first estimated the study specific linear trends between exposure and outcome using the method described by Greenland and Longnecker, which accounts for correlation of the rate ratios within each study to avoid potential bias. The estimated linear trends were then pooled with random-effects meta-analysis. Next, we explored potential non-linear dose-response relation in each study by using restricted cubic splines with three knots in the dose-response regression model, and results from each study were then pooled together with random effects multivariate meta-analysis. The linear and non-linear models were compared with likelihood ratio tests. We also examined sex specific effects by conducting separate meta-analyses for men and women.

For sensitivity analysis, we first repeated the analysis after excluding rate ratios that had an assigned fasting blood glucose concentration over 7.0 mmol/L, the cut off for diagnosing diabetes, to assess whether the observed trend was mainly attributable to an increased rate of pancreatic cancer associated with type 2 diabetes and to exclude reverse causality from diabetes induced by pancreatic cancer. Secondly, we repeated the analysis after excluding studies that were incorporated using estimated fasting blood glucose concentrations or those that measured fasting blood glucose with variable fasting time among participants. Lastly, we conducted sensitivity analysis to examine whether using minimally adjusted rate ratios from included studies or different methods to assign blood glucose concentrations for open ended categories (the upper bound minus 7.0 mmol/L) would modify the results.
a half or the whole width of the neighboring category for the lowest category, and the lower bound plus the width or twice the width of the neighboring category for the highest category) would influence our results. The significance level was set at 5% throughout this study. The statistical analyses were performed with Stata version 12 (StataCorp, TX).

Results

We identified 2769 articles for review of title and abstract (fig 1). After the initial screening, full text of potentially eligible articles was retrieved for detailed assessment. Nine eligible studies were included for meta-analysis, with a total of 2408 patients with pancreatic cancer. Table 1 summarises included studies, and table 2 shows rate ratios for various glucose concentrations. All included studies have been published as full manuscripts and are of high quality (see fig A, appendix 2).

Five studies used fasting blood glucose as the exposure. Among them, fasting time was variable in one study, with 58.0% of the participants fasted for more than 4 hours. Fasting blood glucose was derived from haemoglobin A1c in two studies and from post-load blood glucose concentration in two studies. We calculated the number of cases of pancreatic cancer in each glucose category using the number of study population and cumulative incidence in one study. Seven studies reported hazard ratios for pancreatic cancer. Odds ratios were reported in two nested case control studies, which used incidence density sampling for selecting control subjects; therefore, the obtained odds ratios mean rate ratios. Three studies exclusively reported sex specific rate ratios. In two of them we carried out pooling of sex specific rate ratios for subsequent meta-analysis, but this was not feasible in one study because the categorisation of blood glucose was different between sexes. We used sex specific rate ratios to incorporate that study.

For high versus low meta-analysis, we included all nine studies. No significant small study bias was found (Begg’s test P=0.59, Egger’s test P=0.71; fig B in appendix 2). The pooled rate ratio of pancreatic cancer for the highest versus the lowest category of fasting blood glucose was 1.83 (95% confidence interval 1.50 to 2.24), with significant heterogeneity (I^2=52.4%, P=0.026) (fig 2).

For dose-response meta-analysis, we excluded one study that divided glucose concentration into only two categories because at least three exposure categories are needed to estimate the study specific trend. Figure 3 summarises the estimated study specific linear trends of the relation between fasting blood glucose and rate ratio and the pooled estimate from random effects meta-analysis. There was a positive dose-response relation between fasting blood glucose concentration and the rate of pancreatic cancer (pooled rate ratio was 1.14 (95% confidence interval 1.06 to 1.24) per 0.56 mmol/L increase in fasting blood glucose concentration), without significant heterogeneity across studies (P=0.21, I^2=26.8%). No individual study had excessive influence on the summary estimate (see fig C in appendix 2).

Figure 4 shows the results of non-linear dose-response meta-analysis. There was a linear relation between fasting blood glucose and the rate of pancreatic cancer across both prediabetes and diabetes, whereas no significant non-linear association was noted (P=0.52 for the comparison between linear and non-linear models). The pooled rate ratio of pancreatic cancer per 0.56 mmol/L increase in fasting blood glucose was 1.14 (95% confidence interval 1.06 to 1.22, P=0.001) across the range between 4.1 mmol/L and 10.6 mmol/L, without significant heterogeneity. For sex specific analyses, five studies reported rate ratios for men and three studies reported rate ratios for women. The results of sex specific analyses (see appendix 2 for meta-analyses of high v low (fig D), linear trend (fig E), and non-linear dose-response in men/women (figs F and G)) were in line with those of analyses combining both sexes, showing a linear relation between fasting blood glucose and the rate of pancreatic cancer. The pooled rate ratio of pancreatic cancer per 0.56 mmol/L increase in fasting blood glucose was 1.05 (0.99 to 1.12; P=0.13) in men and 1.17 (1.07 to 1.29; P=0.001) in women.

Sensitivity analyses showed similar results. When we excluded categories with an assigned fasting blood glucose concentration over 7.0 mmol/L, the pooled rate ratio per 0.56 mmol/L increase in fasting blood glucose was 1.15 (95% confidence interval 1.05 to 1.27; P=0.003) (fig 5). After we excluded two studies that measured post-load blood glucose and two studies that measured haemoglobin A1c the pooled rate ratio per 0.56 mmol/L increase in fasting blood glucose was 1.11 (1.02 to 1.20; P=0.021) (see fig H, appendix 2). Exclusion of the study that had variable fasting time or use of minimally adjusted rate ratios for meta-analysis had little influence on the results (pooled rate ratios per 0.56 mmol/L increase in fasting blood glucose were 1.14 (1.07 to 1.21; P<0.001) and 1.13 (1.07 to 1.21; P<0.001), respectively). Use of different methods to assign blood glucose concentrations for open ended categories also had negligible impacts on the results (see table A in appendix 2).

Discussion

There is a linear dose-response relation between fasting blood glucose concentration and the rate of pancreatic cancer across prediabetes and diabetes, with every 0.56 mmol/L increase in fasting blood glucose associated with a 14% increase in the rate of pancreatic cancer. We used meta-analysis to evaluate the association between prediabetes and pancreatic cancer and to investigate the dose-response relations between blood glucose concentrations and risk of pancreatic cancer.

Comparisons with individual studies

This meta-analysis provides novel insights into the association between abnormal glucose metabolism and pancreatic cancer. In previous research, the association between prediabetes and pancreatic cancer was significant in only one study but non-significant in four, mixed in three, and unclear in one study. By pooling all those studies, this meta-analysis shows that the rate of pancreatic cancer increases linearly with worsening hyperglycaemia throughout prediabetes and diabetes. Several reasons might explain the seemingly conflicting results of those individual studies. The distribution and categorisation of glucose concentration varied among studies, thus the ranges of glucose concentrations compared to derive the rate ratios differed among studies and the rate ratios could not be directly compared. Secondly, many categories of glucose concentration in the range of prediabetes contained only a small number of patients with pancreatic cancer, yielding rate ratios with wide 95% confidence intervals (table 2). By contrast, our meta-analysis assessed the change in the rate of pancreatic cancer per unit increase in fasting blood glucose in each study, and pooling of multiple studies provided greater statistical power and more precise estimates. It is worth noting that while the rate ratio for the highest versus the lowest blood glucose categories was heterogeneous among included studies, the trend of rate ratio in relation to blood glucose was actually similar across
those studies without significant heterogeneity. Our results support that fasting hyperglycaemia is a dose-dependent risk factor for pancreatic cancer.

Possible explanations and implications

The dose-response relation between blood glucose concentration and risk of pancreatic cancer might be attributed to the fact that pancreatic cancer cells depend heavily on glucose for growth. While the normal pancreas metabolises glucose through oxidative phosphorylation, pancreatic cancer cells preferentially metabolise glucose through aerobic glycolysis, which generates less energy but more metabolites required for biosynthetic functions to sustain cell proliferation and thus confers a survival advantage (Warburg effect). To compensate for the inefficient energy production and meet the growing need for energy and biosynthesis, pancreatic cancer cells have a high requirement for glucose ("glucose addiction") and exhibit increased glucose uptake. Therefore, hyperglycaemia might increase the risk of pancreatic cancer by providing more glucose to fuel tumor growth.

Our results support that prediabetes/type 2 diabetes is a modifiable risk factor for pancreatic cancer in addition to smoking and obesity. Although we cannot completely exclude the possibility of residual confounding by smoking and obesity, the potential extent of confounding is likely small. In all but one study included for dose-response meta-analysis, the rate ratios had been adjusted for smoking and body mass index (BMI). While Jee and colleagues presented rate ratios that adjusted for smoking but not BMI, they pointed out that their participants were far leaner than those in studies from Western populations and further adjustment for BMI had little influence on the rate ratios.

The discovery that risk of pancreatic cancer progressively increases with worsening hyperglycaemia has important implications. Our findings imply that the increasing incidence of pancreatic cancer might be attributed to the rapid increase in prediabetes/diabetes, a global epidemic affecting 14.2% (629 million) of the world’s adult population. Furthermore, prediabetes could provide an important opportunity for prevention of pancreatic cancer, the most effective strategy to reduce related mortality given that pancreatic cancer evades early detection and responds poorly to treatment. Prediabetes precedes overt type 2 diabetes and can be improved or even reversed through changes in lifestyle. Two previous randomised trials have shown that counselling on weight loss, diet, and physical activity for individuals with prediabetes could decrease fasting blood glucose concentrations and reduce the risk of progression to type 2 diabetes by about 60%. Efforts toward early detection of prediabetes in conjunction with implementation of lifestyle changes to improve glucose metabolism could represent a viable strategy to curb the increasing incidence of pancreatic cancer and should be further evaluated.

Strengths and limitations of this study

Our study deals with the limitations of existing research and has several strengths. Previous studies assumed a linear increase in risk with rising blood glucose from normal to diabetes, without assessing whether a threshold for blood glucose concentration existed and if the observed trend resulted mainly from increased risk in the range of diabetes. Furthermore, pancreatic cancer induces diabetes in about 40% of patients, raising concerns that the association between diabetes and pancreatic cancer could be partly because of reverse causality. To resolve these issues, we examined both linear and non-linear dose-response relations between fasting blood glucose and risk of pancreatic cancer to assess whether there is a threshold for blood glucose concentration, rather than assuming linearity without justification. To confirm the association between prediabetes and pancreatic cancer and to exclude possible reverse causality from diabetes induced by pancreatic cancer, we conducted pre-specified sensitivity analysis excluding glucose categories in the diabetes range. We also accounted for correlation among the rate ratios based on a common reference group in each study to avoid underestimating the variance of the study specific trend, which could yield misleading results.

Our results are robust and provide firm support that prediabetes is also associated with increased risk of pancreatic cancer.

Our study also has some limitations. Firstly, only a few studies reported sex specific rate ratios; therefore, results from sex specific meta-analyses were based on a small amount of evidence and should be interpreted with caution. Secondly, information on the use of antidiabetic drugs was not available in the included studies. During follow-up, participants with diabetes might receive antidiabetic drugs that could cause changes in blood glucose concentrations from baseline.

Participants without diabetes, however, were unlikely to receive antidiabetic drugs, and this factor should have little influence on our results in the blood glucose range below diabetes. Thirdly, we excluded one study that reported rate ratios of pancreatic cancer for diabetes and prediabetes because it did not provide the distribution of blood glucose concentration and the number of cases in each category, and thus we could not determine the blood glucose concentrations associated with the rate ratios with precision and estimate the correlation among the rate ratios. When we included this study in the analysis with imputed blood glucose concentrations, and assumed no correlation among the rate ratios, there was little influence on our results. We also excluded one study that did not categorise blood glucose but reported the linear trend between fasting blood glucose and rate ratio of pancreatic cancer. Although this study could be incorporated into the linear dose-response meta-analysis, we decided to completely exclude this study because blood glucose was measured after diagnosis of pancreatic cancer in 19 of its 48 patients, raising concerns of reverse causality from diabetes induced by pancreatic cancer. Lastly, our meta-analysis was conducted with summary statistics rather than individual data. Access to and examination of data from individual participants could allow more precise delineation of the dose-response relation and further control of potential residual confounding.

Conclusions

Our dose-response meta-analysis shows that every 0.56 mmol/L increase in fasting blood glucose is associated with a 14% increase in the rate of pancreatic cancer. Prediabetes is also a risk factor for pancreatic cancer and provides an opportunity for prevention of pancreatic cancer.

Contributors: All authors had full access to all of the data in the study and take responsibility for the integrity of the data and the accuracy of the data analysis. W-CL, Y-KT, and K-LC contributed to study concept and design, data analysis and interpretation, and writing of the report. M-SW, J-TL, and H-PW contributed to data interpretation and provided intellectual enrichment to the report. All authors met the ICMJE criteria for authorship. All authors agree with the manuscript results and conclusions. Y-KT and K-LC are guarantors.

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What is already known on this topic

Type 2 diabetes is an established risk factor for pancreatic adenocarcinoma, the most lethal cancer

Prediabetes precedes type 2 diabetes and can be improved or reversed through lifestyle changes, suggesting that prediabetes might be a risk factor for pancreatic cancer and an opportunity for intervention.

Whether prediabetes increases the risk of pancreatic cancer remains unclear; previous studies yielded inconsistent results, and no systematic review has assessed the dose-response relation between blood glucose and risk

What this study adds

The rate of pancreatic cancer increases linearly by 14% with every 0.56 mmol/L (10 mg/dL) increase in fasting glucose across both prediabetes and diabetes.

Efforts toward early detection of prediabetes and lifestyle changes to improve glucose metabolism could represent a viable strategy to curb the increasing incidence of pancreatic cancer

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Transparency: The lead author (the manuscript’s guarantor) affirms that the manuscript is an honest, accurate, and transparent account of the study as planned; if available, registered) have been explained.

Datasharing: No additional data available.

Transparency: The lead author (the manuscript’s guarantor) affirms that the manuscript is an honest, accurate, and transparent account of the study as planned; if available, registered) have been explained.

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### Tables

#### Table 1 | Summary of prospective studies included in systematic review and dose-response meta-analysis on blood glucose concentration and rate of pancreatic cancer

| Study                                      | Region               | Design                 | Mean age (years) | Women (%) | Baseline | Duration of follow-up (year) | Adjusted variables                                      |
|--------------------------------------------|----------------------|------------------------|------------------|-----------|----------|-------------------------------|---------------------------------------------------------|
| Gapstur, 2000†                           | North America        | Cohort                 | 39.9             | 42.6      | 1963-73  | 25 (mean)                     | Age, race, smoking, BMI                                 |
| Batty, 2004**                            | Europe               | Cohort                 | 51.5             | 0         | 1967-70  | 25                            | Age, smoking, BMI, physical activity, etc               |
| Stolzenberg-Solomon, 2005**              | Europe               | Case cohort            | 57.2             | 0         | 1985-88  | 13.8 (median)                 | Age, smoking, BMI                                       |
| Jee, 2005†                               | Asia                 | Cohort                 | 46.9             | 36.1      | 1992-95  | 10                            | Age, age², smoking, alcohol                             |
| Ansary-Moghadam, 2006**                  | Asia, Australia/New Zealand | Cohort               | 46.3             | 35.3      | 1961-99  | 6.8 (median)                  | Age, sex, study, smoking, BMI                           |
| Inoue, 2009†                             | Asia                 | Cohort                 | 55.8             | 65.6      | 1990-94  | 10.2 (mean)                   | Age, area, smoking, alcohol, cholesterol               |
| Johansen, 2010‡                           | Europe               | Cohort                 | Male 43.9        | —         | 1972-2005 | 12.8 (mean)                   | Smoking, BMI, age                                      |
|                                           |                      |                        | Female 44.1      | 49.9      | 1972-2005 | 11.3 (mean)                   |                                                         |
| Grote, 2011†                             | Europe               | Nested case-control    | 58 (cases), 58 (control) | 51.7 (cases), 51.7 (control) | 1992-2000 | 5.3 (mean) | Smoking, BMI, matched for date, sex, age, food, drink, centre |
| Wolpin, 2013‡                             | North America        | Nested case-control    | 63.1 (cases), 62.5 (control) | 71.5 (cases), 70.7 (control) | 1976-98 | 12.2-25.3 (median) | Cohort, smoking, BMI, fasting time, age, race, sex |

BMI = body mass index.

*Pooled analysis of 30 cohorts. Australia: Busselton, Canberra-Queanbeyan, Long, Study of Aging, Melbourne, National Heart Foundation, Newcastle, Perth, WA AAAScreenees; New Zealand: Fletcher Challenge; China: Anzhen, EastBeijing, GuangzhouOccupational, SevenCitiesCohorts, SixCohorts, Tianjin, Xi’an; Hong Kong: Hong Kong; Japan: Aita town, Akabane, CivilService Workers, Hisayama, Konan, Ohasama, Saitama, Shibata, Shigaraki Town, Shirakawa; Singapore: Singapore Heart; South Korea: KMIC; Taiwan: CVDFACTS.

†Pooled analysis of seven cohorts. Austria: Vorarlberg Health Monitoring and Prevention Program; Norway: Oslo study I cohort, Norwegian Counties Study, Cohort of Norway, Age 40-programme; Sweden: Västerbotten Intervention Project, Malmö Preventive Project.

‡Pooled analysis of five cohorts in US: HealthProfessionals Follow-up Study, Nurses’ Health Study, Physicians’ Health Study, Women’s Health Initiative—Observational Study, Women’s Health Study.
| Study                  | Rate ratio (95% CI) | No of cases/total or person years (PY) |
|------------------------|---------------------|----------------------------------------|
| Gapstur, 2000†         |                     |                                        |
| <5.1*                  | 1.65 (1.05 to 2.60) | 30/379 686 PY                           |
| 5.1-5.9*               | 1.60 (0.95 to 2.70) | 31/116 475 PY                           |
| ≥7.0*                  | 2.15 (1.22 to 3.80) | 23/52 731 PY                            |
| Batty, 2004§           |                     |                                        |
| Men:                   |                     |                                        |
| <5.6*                  | 1.35 (0.66 to 2.80) | 34/133                                  |
| 5.6-6.9*               | 3.99 (1.44 to 11.0) | 10/58 843                               |
| ≥7.0*                  | 3.99 (1.44 to 11.0) | 10/58 843                               |
| Stolzenberg-Solomon, 2005§ |       |                                        |
| Men:                   |                     |                                        |
| <5.2                   | 1.15 (0.86 to 2.59) | 53/137                                  |
| 5.5-5.9                 | 1.69 (0.97 to 2.94) | 32/146                                  |
| ≥5.9                   | 1.69 (0.97 to 2.94) | 32/146                                  |
| Jee, 2005§             |                     |                                        |
| Men:                   |                     |                                        |
| <5.0                   | 1.08 (0.95 to 1.24) | 43/304 362†                             |
| 5.0-6.0                 | 1.34 (1.09 to 1.64) | 10/58 020†                              |
| 7.0-7.7                 | 1.37 (1.01 to 2.00) | 12/11 459†                              |
| ≥7.8                   | 2.09 (1.70 to 2.58) | 12/11 459†                              |
| Women:                 |                     |                                        |
| <5.0                   | 1.27 (1.03 to 1.57) | 15/157 940†                             |
| 5.0-6.0                 | 1.39 (1.06 to 2.02) | 32/22 579†                              |
| 7.0-7.7                 | 1.99 (1.13 to 3.49) | 1/565†                                  |
| ≥7.8                   | 1.67 (1.09 to 2.56) | 1/565†                                  |
| Ansary-Moghaddam, 2006‡ |                     |                                        |
| <5.2                   | 1.79 (1.03 to 3.10) | 28/125 855                               |
| 5.2-5.8                 | 1.79 (1.03 to 3.10) | 28/125 855                               |
| >5.8                   | 2.08 (1.18 to 3.67) | 28/125 855                               |
| Inoue, 2009§           |                     |                                        |
| Men:                   |                     |                                        |
| <5.6                   | 0.74 (0.24 to 2.22) | 4/21 687 PY                              |
| >5.6                   | 1.00 (0.42 to 2.39) | 4/21 687 PY                              |
| Women:                 |                     |                                        |
| <5.6                   | 0.81 (0.60 to 1.08) | 35/165 838 PY                            |
| >5.6                   | 1.14 (0.88 to 1.49) | 35/165 838 PY                            |
| Johansen, 2010§         |                     |                                        |
| Mean (SD) men:         |                     |                                        |
| 4.2 (0.5)              | 1.00 (0.42 to 2.39) | 102/772 727 PY‡                          |
| 4.8 (0.3)              | 0.81 (0.60 to 1.08) | 102/772 727 PY‡                          |
| 5.1 (0.5)              | 1.14 (0.88 to 1.49) | 102/772 727 PY‡                          |
| 5.6 (0.3)              | 1.01 (0.76 to 1.34) | 102/772 727 PY‡                          |
| 6.9 (2.0)              | 1.24 (0.95 to 1.61) | 102/772 727 PY‡                          |
| Mean (SD) women:       |                     |                                        |
Table 2 (continued)

| Fasting blood glucose (mmol/L) | No of cases/total or person years (PY) | Rate ratio (95% CI) |
|--------------------------------|----------------------------------------|--------------------|
| 4.1 (0.6)                      | 34/666 667 PY¶                          | 1                  |
| 4.3-5.3**                      | 72/173                                  | 1                  |
| 5.4-5.8**                      | 131/282                                 | 1.27 (0.84 to 1.93) |
| 5.9-6.1**                      | 102/184                                 | 1.77 (1.14 to 2.75) |
| 6.2-6.9**                      | 97/188                                  | 1.46 (0.93 to 2.30) |
| 7.0-14.2**                     | 54/85                                   | 2.42 (1.33 to 4.39) |
| WOLPIN, 2013††                 |                                        |                    |
| Median 4.2**                   | 61/246                                  | 1                  |
| Median 4.5**                   | 92/276                                  | 1.59 (1.07 to 2.36) |
| Median 4.7**                   | 101/286                                 | 1.62 (1.22 to 2.70) |
| Median 5.0**                   | 74/262                                  | 1.36 (0.89 to 2.07) |
| Median 5.4**                   | 100/285                                 | 1.79 (1.17 to 2.72) |

*Estimated from post-load blood glucose concentration (see method).
†Case number calculated from reported number of study population and cumulative incidence.
‡Pooled analysis of 30 cohorts (see table 1).
§Pooled analysis of seven cohorts (see table 1).
¶Person year calculated from reported number of cases and incidence rate.
**Estimated from haemoglobin A1c (see methods).
††Pooled analysis of five cohorts in US (see table 1).
Figures

**Fig 1** Flow chart of literature search for studies investigating association between blood glucose concentration and risk of pancreatic cancer

**Fig 2** Summary rate ratio of pancreatic cancer, highest v lowest blood glucose category. Weights from random effects analysis. *Estimated from reported post-load blood glucose or haemoglobin A1c concentrations. †Pooled from rate ratios for men and women. ‡Pooling of rate ratios for men and women not feasible because categorisation of blood glucose differed between sexes
**Fig 3** Summary linear trend of rate ratio per 0.56 mmol/L (10 mg/dL) increase in fasting blood glucose. Weights from random effects analysis.

| Author                    | Rate ratio (95% CI) | Weight (%) | Rate ratio (95% CI) |
|---------------------------|---------------------|------------|---------------------|
| Gapstur 2000              | 1.26 (0.99 to 1.60) | 9          | 1.26 (0.99 to 1.60) |
| Batty 2004 (male)         | 1.08 (0.72 to 1.60) | 4          | 1.08 (0.72 to 1.60) |
| Stolzenberg-Solomon 2005 (male smoker) | 1.19 (0.84 to 1.66) | 5          | 1.19 (0.84 to 1.66) |
| Jee 2005                  | 1.06 (1.00 to 1.13) | 36         | 1.06 (1.00 to 1.13) |
| Ansary-Moghaddam 2006     | 1.48 (0.98 to 2.24) | 3          | 1.48 (0.98 to 2.24) |
| Johansen 2010             | 1.02 (0.88 to 1.19) | 17         | 1.02 (0.88 to 1.19) |
| Male                      | 1.30 (1.04 to 1.64) | 10         | 1.30 (1.04 to 1.64) |
| Female                    | 1.23 (1.00 to 1.49) | 12         | 1.23 (1.00 to 1.49) |
| Grote 2011                | 1.49 (0.99 to 2.24) | 4          | 1.49 (0.99 to 2.24) |
| Wolpin 2013               | 1.14 (1.06 to 1.24) | 100        | 1.14 (1.06 to 1.24) |

Overall: $I^2=27\%, p=0.206$

**Fig 4** Dose-response relation between fasting blood glucose and rate ratio for pancreatic cancer, showing point estimates and 95% confidence interval for non-linear analysis and point estimates for linear analysis. Circles indicate adjusted rate ratios in individual studies; size of bubble is proportional to precision (inverse of variance) of rate ratio.

**Fig 5** Dose-response relation between fasting blood glucose and rate ratio for pancreatic cancer, excluding categories with assigned fasting blood glucose concentration >7.0 mmol/L. Graph shows point estimates and 95% confidence interval for non-linear analysis and point estimates for linear analysis. Circles indicate adjusted rate ratios in individual studies; size of bubble is proportional to precision (inverse of variance) of rate ratio.