Epileptiform activity predicts epileptogenesis in cerebral hemorrhage

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Abstract
This retrospective case–controlled study was performed to evaluate whether Epileptiform Activity, suspected clinical seizures, and/or 2HELPS2B/S after nontraumatic Intraparenchymal Hemorrhage or Subarachnoid Hemorrhage can predict Epilepsy. Hundred and thirty-two patients were included—29 (Epilepsy), 103 (Control Group). After matching, the average effect for all three risk factors was significant as follows: (1) Epileptiform Activity ($p = 0.012$, odds ratio 3.14), (2) suspected seizures ($p = 0.021$, odds ratio 3.78), and (3) 2HELPS2B/S score ($p < 0.001$, odds ratio 4.94). This study shows that Epileptiform Activity, suspected seizures, and particularly, the 2HELPS2B/S score in the acute phase are risk factors for the development of epilepsy after nontraumatic intraparenchymal and subarachnoid hemorrhage.

Introduction
Brain injury is a common cause of epilepsy and results in decreased quality of life and increased mortality.¹ The latent period between brain injury and epilepsy is a target for antiepileptogenesis therapy. Postinjury epilepsy is an attractive research area as it has animal model analogs, is a capturable population for trials, and has a window for intervention. Despite success in animal models, the clinical trial results have been disappointing.² One element that may improve future trials is enriching the study population with those most likely to develop epilepsy. Developing biomarkers that can risk stratifying postinjury patients has evident clinical utility and can improve clinical trial efficacy.³

One potential biomarker is electroencephalogram (EEG) Epileptiform Activity. These EEG abnormalities have long been recognized as a harbinger of epilepsy in animal models⁴ and were recently found to be associated with an increased risk of posttraumatic epilepsy.⁵ Similarly, acute symptomatic seizures are also associated with an increased risk of epilepsy after ischemic stroke.⁶ Here we expand upon these findings in a case–controlled study in 132 patients with either subarachnoid hemorrhage (SAH) or intraparenchymal hemorrhage (IPH) with covariate adjustment with coarsened exact matching to determine if epileptiform abnormalities, suspected acute symptomatic seizures, and a seizure risk score—2HELPS2B/S⁷ can predict the development of epilepsy. Further, we examine the influence of antiseizure medications (ASM) on the risk of developing epilepsy.

Methods
This retrospective case–controlled study was approved by the institutional review board (IRB) at the University of Wisconsin—Madison. Waiver for informed consent was granted. Medical records were reviewed for adult patients presenting between 2003 and 2018 with either subarachnoid hemorrhage or intraparenchymal hemorrhage. Most of the patients were identified via our Clinical Research Data Service Query (CDRS) which utilized an algorithmic search of billing codes with the following criteria: (1) patient has an inpatient admission with a final hospital billing diagnosis that includes SAH (ICD codes 430, I60, I69) or intracerebral hemorrhage (ICD codes 431, I61), (2) patient has a billing code for EEG associated with the admission, (3) discharge date within the time frame January 1, 2010 through September 30, 2018, and (4) the patient is 18 years or older on the discharge date.
initial CRDS algorithm yielded 513 patient encounters. Prior to 2010, a separate data collection strategy was employed using the Epic “My Reports” tool using parameters that identified patients who had an EEG and were admitted by the Neurosurgeons at our institution. Both sets of patients were then screened for inclusion and exclusion criteria. Study inclusion criteria included (1) age ≥ 18, (2) nontraumatic SAH or IPH at presentation, (3) routine or continuous EEG monitoring during initial hospital admission within 14 days of presentation, (4) for controls—a clinic follow-up 2 years following the presentation, where it is determined that a patient did not develop epilepsy. Study exclusion criteria included (1) known diagnosis of seizures/Epilepsy, (2) presentation with subdural hematomas, (3) cause of hemorrhage is a conversion from ischemic stroke, (4) recent trauma/surgery, known intracranial neoplasm. After combining both sets of patients, 132 patients were included in the study (Fig. 1). Subjects are considered to have developed Epilepsy if they have at least one documented clinical seizure between 1 and 24 months. The remaining patients were identified as controls. Given the above criteria, all the participants in this study had a full data set except for four controls with missing ASM data.

Three risk factors and ASM use were selected a priori and analyzed to assess their influence on epilepsy development. Epileptiform activity was defined as the presence of EEG markers of cortical irritability within 2-week postinjury. These EEG markers include sporadic epileptiform discharges, seizures, brief potentially ictal rhythmic discharges (BRDs), Lateralized Periodic Discharges (LPDs), Generalized Periodic discharges (GPDs) if > 2 Hz or plus features, and Lateralized Rhythmic Delta Activity (LRDA). The second risk factor was suspected acute symptomatic seizure at presentation. The third risk factor was 2 HELPS 2B/S. 2 HELPS 2B/S is a modified form of 2 HELPS 2B. The difference is that either BRDs or electrographic seizures will net two points as the neurophysiology of BRDs and electrographic seizures are similar. The purpose of 2 HELPS 2B/S was to quantify cortical irritability. ASM variable was defined as being on an ASM either at the time of the first unprovoked epileptic seizure (for subjects) or use of ASM at a 2-year follow-up (for controls).

Univariate comparison prior to matching was performed on risk factors and covariates (Table 1). Statistical analysis was performed with permutation testing. Commonly used clinical break points for continuous variables (intraparenchymal hemorrhage volume) or clinical ordinal scales were defined a priori and used to dichotomize variables.

Using a case-controlled study design, the distribution across selected covariates was adjusted by coarsened exact matching to estimate the marginal odds ratio. Coarsened exact matching overcomes bias introduced with propensity score with greater flexibility than exact matching. The mean difference in covariates and effective sample size was examined after each matching to assess quality, performed with Matchit. Covariates adjusted via matching were selected a priori based on the likelihood of
affecting the risk factor and the development of epilepsy. The number of matched covariates was minimized and dichotomized along standard clinical break points. Highly correlated variables were excluded (e.g., Hunt-Hess and GCS). Credible intervals were calculated with bootstrapping with replacement to the full cohort size. There was no forced balancing between diagnoses. A moderation analysis was performed to assess the potential influence of ASM on postinjury epilepsy.

The classifier ability of 2HELPS2B/S was analyzed with the area under the ROC curve (AUC). All analysis was performed in R version 4.1.1 (R Foundation Vienna, Austria). The P value threshold for significance was set at <0.05. For the marginal odds ratio of the three risk factors (primary outcome), a family-wise error rate was adjusted with the Bonferroni method.

### Results

The final analysis included 132 patients (29 developed epilepsy). Clinical covariates were not statistically different in those who developed epilepsy. However, several EEG factors either had a trend or were statistically significant. The three risk factors of interest were significant: Epileptiform Activity \( p = 0.03 \), Suspected Acute Symptomatic Clinical Seizure \( p = 0.042 \), and 2HELPS2B/S>0<0.001. (Table 1).

### Table 1. Unmatched patient characteristics and Matched Risk Factor analysis.

| Unmatched patient characteristics—n (%) | No Epilepsy | Epilepsy | \( p \) value |
|----------------------------------------|-------------|----------|--------------|
| Demographics                           |             |          |              |
| Age in years (mean)                    | 58.3        | 54.7     | 0.23         |
| Age ≥65 years                          | 34 (33)     | 8 (27)   | 0.74         |
| Sex (76 female, 57.6%)                 | 59 (57)     | 17 (59)  | 1            |
| IPH and SAH (N = 132)                  | 103 (100)   | 29 (100) |              |
| IPH (N = 54)                           | 41 (40)     | 13 (45)  |              |
| SAH (N = 78)                           | 62 (60)     | 16 (55)  |              |
|Clinical factors                        |             |          |              |
| IPH versus SAH (% with IPH)            | 41 (40)     | 13 (48)  | 0.55         |
| ICH score ≥2                           | 8 (20)      | 3 (23)   | 0.85         |
| HH >2                                  | 19 (31)     | 8 (50)   | 0.32         |
| Fisher >2                              | 53 (86)     | 16 (100)| 0.24         |
| Intraventricular hemorrhage            | 42 (68)     | 11 (69)  | 1            |
| Hydrocephalus                          | 41 (66)     | 14 (88)  | 0.17         |
| GCS (mean)                             | 11.1        | 10.4     | 0.44         |
| GCS <8                                 | 31 (30)     | 11 (38)  | 0.57         |
| ASM at time of seizure or 2-year follow-up | 28 (27) | 14 (48) | 0.058       |
| EEG Findings                           |             |          |              |
| Day of the first EEG (mean)            | 4.1         | 3.3      | 0.29         |
| EEG duration hours (mean)              | 23.06       | 47.7     | 0.007        |
| Electrographic seizure                 | 3 (3)       | 4 (14)   | 0.066        |
| BIRD                                   | 0           | 0        | 1            |
| Sporadic epileptiform discharge        | 16 (16)     | 14 (48)  | 0.001        |
| Lateralized periodic discharges        | 3 (3)       | 3 (10)   | 0.23         |
| Lateralized rhythmic delta activity    | 4 (4)       | 4 (14)   | 0.13         |
| Generalized rhythmic delta activity    | 9 (9)       | 4 (14)   | 0.65         |
| Focal slowing                          | 51 (50)     | 16 (55)  | 0.74         |
| Generalized slowing                    | 85 (83)     | 24 (83)  | 1            |
| Burst suppression                      | 9 (9)       | 4 (14)   | 0.66         |
| 2HELPS2B/S Score (mean)                | 0.4         | 1.34     | <0.001       |
| Time to first epileptic seizure (median, IQR) | 318, 214 |              |              |
| Primary risk factors of interest       |             |          |              |
| 2HELPS2B/S>0                           | 30 (29)     | 21 (72)  | <0.001*      |
| Acute symptomatic clinical seizure     | 13 (13)     | 10 (35)  | 0.042*       |
| Any evidence of EEG cortical irritability | 22 (21) | 16 (55) | 0.003*       |

### Table 1. Continued.

| Matched risk factor analysis | Average effect estimate | OR | Credible intervals | \( p \) value |
|-----------------------------|-------------------------|----|--------------------|--------------|
| ASM moderation odds         |                         |    |                    | 0.016        |
| Epileptiform activity with ASM | 2.08       | 0.25–9.57 |             |              |
| Epileptiform activity without ASM | 3.85     | 0.62–25.95 |             |              |
| Marginal Odds               |                         |    |                    |              |
| Suspected acute symptomatic seizure with ASM | 3.79     | 1.00–14.10 | 0.021*      |
| Suspected acute symptomatic seizure without ASM | 0.92     | 0.15–2.30 | 0.17        |
| Suspected acute symptomatic seizure with ASM | 3.1       | 0.00–29.63 |             |              |
| ASM Moderation Odds         |                         |    |                    | <0.001*      |
| 2HELPS2B/S>0 with ASM       | 4.94                   | 1.72–16.64 |             |              |
| 2HELPS2B/S>0 without ASM    | 4.53                   | 0.63–36.83 |             |              |

Abbreviations: ASM, antiseizure medications; EEG, Electroencephalogram; IPH, intraparenchymal hemorrhage; OR, odds ratio; SAH, subarachnoid hemorrhage.

*Adjusted for three multiple comparisons.
After matching, the marginal treatment effect of prolonged ASM use on the development of epilepsy was not significant ($p = 0.36$). The marginal odds ratio for all three risk factors were significant: *Epileptiform Activity* odds ratio = 3.14, $p = 0.012$; *Suspected Acute Symptomatic Clinical Seizure* odds ratio 3.78, $p = 0.021$; 2HELPS2B/S odds ratio 4.94, $p < 0.001$ (Fig. 2, Table 1). Matching details are in Supplement A.

The 2HELPS2B/S score was assessed as a risk factor for the development of postinjury epilepsy in the whole cohort of 132 patients (Fig. 2B-C) with an AUC of 0.725. Subgroup analysis indicated that there was no difference for 2HELPS2B/S as a predictor of Epilepsy between IPH and SAH groups.

**Discussion**

The primary result is that Epileptiform Activity in the acute phase of nontraumatic cerebral hemorrhage is associated with an increased risk of epilepsy. Matching of covariates was performed to minimize the effect of confounders such as hemorrhage size. Additionally, prolonged use of ASM after the acute phase of hemorrhage did not affect epilepsy development. Both Epileptiform Activity and acute symptomatic seizures were significantly associated. The seizure prediction score 2HELPS2B was slightly modified to include electrographic seizures and named 2HELPS2B/S. 2HELPS2B/S combines the other two risk factors and was more powerful than either.

These results are consistent with recent studies. Kim et al. (2018) showed that epileptiform discharges increase the risk of epilepsy after traumatic brain injury. A large study by Galovic et al. (2018) found acute symptomatic seizures to be the best epilepsy predictor following ischemic stroke. Similar to the pivotal clinical trial, ASM did not prevent postbrain injury epilepsy. It remains unclear if the association between Epileptiform Activity in the acute phase of brain injury and the development of epilepsy is causal or related to a similar underlying common cause. Arguing against the hypothesis of causation are the failures of ASM to prevent epilepsy. However, these trials did not identify only the brain-injured patients with acute cortical irritability nor did they titrate ASM based on individual response to therapy. Future trials looking at antiepileptogenesis therapies should include acute phase EEG monitoring to examine these hypotheses. Even if no causal relationship exists, Epileptiform Activity can still be useful for risk stratification purposes in clinical trials.

Limitations include selection bias, as only patients with 2 years of clinical follow-up and an EEG were included. It is possible that the requirement for 2 years of follow-up would disproportionately exclude patients based on insurance, rare, or socioeconomic status. However, we do...
not have evidence that these differences are related to the development of Epilepsy, and further studies are needed to evaluate these links. An ongoing prospective study (EpiBioSRX15) can test these results in a prospective observational study. Matching covariates for case-controlled studies is an imperfect process that does not completely suppress bias; hence, the need for replication. ASM use was also considered a single dichotomized variable due to limits in study size, so an ASM effect based on a dose or a drug mechanism remains possible.

In summary, Epileptiform Activity, suspected acute symptomatic seizures, and particularly the 2HELPS2B/S are promising predictors of postbrain injury epilepsy and should be validated in large prospective observational studies.

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Authors’ Contribution

Tseun Han James Kong is the first author of this paper and was involved in all aspects of this project except for the statistical analysis itself. Mohammad Abdul Azeem, Ayesha Naem, and Shawn Allen were directly involved in the acquisition of data as well as the review and editing of this manuscript. Jennifer Ahjin Kim and Aaron F Struck played a vital role in the conception/design of the study and in assisting in drafting this manuscript. Dr Struck was also responsible for performing the statistical analysis.

Conflict of Interest

No relevant conflict of interest is present for the authors. This study is not a clinical trial and does not use material reproduced from other sources.

Data Availability Statement

Data are available on request due to ethical considerations.

References

1. Englander J, Bushnik T, Wright JM, Jamison L, Duong TT. Mortality in late post-traumatic seizures. J Neurotrauma. 2009;26(9):1471-1477. doi:10.1089/neu.2008.0729
2. Temkin NR, Dikmen SS, Wilensky AJ, Keihm J, Chabal S, Winn HR. A randomized, double-blind study of phenytoin for the prevention of post-traumatic seizures. N Engl J Med. 1990;323(8):497-502. doi:10.1056/NEJM199008233230801
3. Goldenholz DM, Tharayil J, Moss R, Myers E, Theodore WH. Monte Carlo simulations of randomized clinical trials in epilepsy. Ann Clin Transl Neurol. 2017;4(8):544-552. doi:10.1002/acn3.426
4. Staley KJ, White A, Dudek FE. Interictal spikes: harbingers or causes of epilepsy? Neurosci Lett. 2011;497(3):247-250. doi:10.1016/j.neulet.2011.03.070
5. Kim JA, Boyle EJ, Wu AC, et al. Epileptiform activity in traumatic brain injury predicts post-traumatic epilepsy. Ann Neurol. 2018;83(4):858-862. doi:10.1002/ana.25211
6. Galovic M, Dohler N, Erdelyi-Canavese B, et al. Prediction of late seizures after ischaemic stroke with a novel prognostic model (the SeLECT score): a multivariable prediction model development and validation study. Lancet Neurol. 2018;17(2):143-152. doi:10.1016/S1474-4422(17)30404-0
7. Struck AF, Ustun B, Ruiz AR, et al. Association of an Electroencephalography-Based Risk Score with Seizure Probability in hospitalized patients. JAMA Neurol. 2017;74(12):1419-1424. doi:10.1001/jamaneurol.2017.2459
8. Yoo JY, Jett C, Kwon CS, et al. Brief potentially ictal rhythmic discharges and paroxysmal fast activity as scalp electroencephalographic biomarkers of seizure activity and seizure onset zone. Epilepsia. 2021;62(3):742-751. doi:10.1111/epi.16822
9. Iacus SM, King G, Porro G. Causal inference without balance checking: coarsened exact matching. Polit Anal. 2012;20(1):1-24. doi:10.1093/pan/mpr013
10. King G, Nielsen R. Why propensity scores should not be used for matching. Political Analysis. 2019;27(4): 435-454.
11. Ho D, Imai K, King G, Stuart EA. MatchIt: nonparametric preprocessing for parametric causal inference. J Stat Softw. 2011;42:1-28. doi:10.18637/jss.v042.i08
12. Jung K, Lee J, Gupta V, Cho G. Comparison of bootstrap confidence interval methods for GSCA using a Monte Carlo simulation. Front Psychol. 2019;10:2215. doi:10.3389/fpsyg.2019.02215
13. Green KM, Stuart EA. Examining moderation analyses in propensity score methods: application to depression and substance use. J Consult Clin Psychol. 2014;82(5):773-783. doi:10.1037/a0036515
14. Robin X, Turck N, Hainard A, et al. pROC: an open-source package for R and S+ to analyze and compare ROC curves. BMC Bioinformatics. 2011;12(1):77. doi:10.1186/1471-2105-12-77
15. Vespa PM, Shrestha V, Abend N, et al. The epilepsy bioinformatics study for anti-epileptogenic therapy (EpiBioS4Rx) clinical biomarker: study design and protocol. Neurobiol Dis. 2019;123:110-114. doi:10.1016/j.nbd.2018.07.025

Supporting Information
Additional supporting information may be found online in the Supporting Information section at the end of the article.

Supplement A. Matching Process - Summary of balance for all data with coarsened exact matching.