Examining the Threshold Effect of Exchange Rate Changes on Monetary Policy Reaction Function of ASEAN-5: A Panel Threshold Approach

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Abstract

In small open economies, the connection between exchange rate and monetary policy function is well-recognised either in the form of theoretical concept or empirical literature. Unlike others, this paper revisits the study by utilising the panel threshold approach of Hansen (1999) to investigate how the monetary policy function of ASEAN-5 responding to exchange rate changes in two different policy regimes, pre-crisis (1980Q1-1996Q4) and post-crisis (1999Q1-2015Q4). The results exhibit asymmetric effect of exchange rate changes on monetary policy function in both sub-periods. The double threshold effect in the pre-crisis, showing exchange rate changes influencing the policy function in a limited threshold value. Conversely, in the post-crisis, the single threshold effect illustrating exchange rate changes having stronger effect on the policy function in various magnitudes at all threshold values. Although claiming the execution of flexible exchange rate system after the crisis, there exist policymakers’ intervention in stabilising the exchange rate changes with respect to ‘fear of floating’ behaviour in ASEAN-5. Lastly, in each sub-period, the threshold effect of exchange rate changes relative to inflation variation is highly significant while trivial to output gap in triggering the policy function. This witnessing, by and large, the ultimate goal of ASEAN-5 is to achieve price stability.

Keywords: Panel threshold; Asymmetric effect; Fear of floating.

1. Introduction

In a small and open emerging country, the vulnerability of such economy towards external shocks especially to exchange rate changes is one of the common characteristics (Mohanty and Klau, 2004). This is for two crucial reasons. Firstly, from the liability dollarization effect on the balance sheet, as most emerging countries possess their debts in foreign currency but the assets are within their local currency. Secondly, from the price level effect, where pass-through from the exchange rate changes into inflation is much greater in emerging economies (Devereux and Lane, 2006). On the surface, these are the leading factors causing most policymakers to intervene in restricting their exchange rate changes even officially declaring the execution of flexible exchange rate system. This is theoretically termed as ‘fear of floating’ behaviour (Calvo and Reinhart, 2002) which can be noticed empirically when the monetary policy reaction function is alternatively reacting to exchange rate changes instead of focusing on other macroeconomic variables (Mohanty and Klau, 2004). Hence, it is interesting to investigate the response of monetary policy reaction function based on exchange rate changes in a small open economy situation.

Therefore, in the present study, we focused on small open economy of ASEAN-5 by observing the behaviour of monetary policy reaction function from the effect of exchange rate changes. For this, the panel threshold method of Hansen (1999) is executed to compare the results between two policy regimes, the pre- and post-crisis periods. Since the majority of these countries have switched from a rigid (fixed) exchange rate to a flexible (floating) exchange rate regime accompanied with inflation targeting policy after the financial crisis of 1997-98. We further extended the studies by capturing the asymmetric effects (if any) of the exchange rate changes (external shock) on monetary policy reaction function. Then, we determine the threshold level of exchange rate changes which trigger significant reaction in the policy function. Finally, we also seek to compare the threshold effect of exchange rate changes relative to inflation variation and the output gap in stimulating the monetary policy reaction function.

This study deviates from past research in two aspects. Firstly, we use the panel threshold approach of Hansen (1999), which seems not to have applied in this area of studies. Secondly, the existing research to examine the asymmetric effect of exchange rate changes on monetary policy reaction function in ASEAN-5 are with whole sample data-set. Conversely, in this research, we focus on two-panel data-sets by comparing the asymmetric effect of exchange rate changes on policy function between pre- and post-crisis periods as of the major shift in the monetary policy regime thereafter the financial crisis (1997-98).

By fixing the threshold variable as exchange rate changes (DLREER), the results from pre- and post-crisis periods proving the influences of exchange rate changes on monetary policy reaction function. In a deeper manner, through observing the unequal values of the coefficient in each threshold value. We also rectify the existence of asymmetric effect of exchange rate changes on monetary policy function in both policy regimes. In the case of pre-crisis, the double threshold effect displays that exchange rate changes only impact the policy function in a restricted...
interval (-0.0070 < \text{DLREER} \leq 0.0126). This reflects the fixed exchange rate system of ASEAN-5. In contrast, using a single threshold affect the exchange rate changes have multiple impacts on the monetary policy in the post-crisis period. Where, when \text{DLREER} \leq -0.0231, policy rate is reduced to cater appreciation of local currency. While, when \text{DLREER} > -0.0231, policy rate is increased to accommodate depreciation of the local currency. Next, observing the threshold effect of exchange rate changes in two different angles either relative to inflation variation or relative to output gap in triggering the monetary policy function. The results in both regimes illustrating that the policy function is strongly reactive to inflation variation while not all to the output gap.

The remainder of this paper proceeds as follows. Section two presents the literature review. Section three is about data, model, and method. Section four interprets the results and findings. Finally, section five concludes.

2. Literature Review
2.1. Reflection – Monetary policy of ASEAN-5

Prior to Asian financial crisis 1997-98, ASEAN-5 was well-known with continuous outperforming economic growth which was motivated by a market-friendly ideology and outward-oriented liberalised foreign policies (Muniandy and Uning, 2006); (Quan, 1998). The financial crisis attack in July 1997, depreciated all ASEAN-5 currencies which started from Thailand and later extended to Malaysia, Indonesia, Philippines as well as Singapore (Mohammad, 2006). In such the maximum level of depreciation was felt by Indonesian rupiah and the minimum depreciation was on Singaporean dollar (Lloyd and Smith, 2004).

In the wake of the crisis, several transformations, especially in the monetary policy setup and financial reformulation, remarkably implemented. In order to achieve price stability and sustainable output growth, adoption of inflation targeting monetary policy framework became essential (Chang and Jong, 2002). In such Indonesia and Thailand adopted inflation targeting policy in January and April 2000 respectively, while the Philippines adopted inflation targeting in January 2002 (Ito, 2010). Concurrently, all the ASEAN-5 countries except Malaysia (MYR-USD pegged) shifted towards greater flexibility in their exchange rate system (Muniandy and Uning, 2006). Besides that, the ASEAN-5 also experienced greater accumulation of foreign reserves. This coincides with persistent current account surplus in the majority of these countries (Eliza et al., 2008).

Although exchange rate system was not the sole factor that instigated the financial turmoil in ASEAN-5. Yet, considering exchange rate variation in policymakers’ decision making is still being crucial as this variable accounts in determining the economic performance of ASEAN-5 (Muniandy and Uning, 2006); (Rana, 1998).

2.2. Empirical Results – Effect of Exchange Rate Changes on Monetary Policy Reaction Function

Most empirical studies supervised on small open economies recorded significant influences of exchange rate on the monetary policy reaction function. To address a few, studies conducted by Mohanty and Klau (2004) on thirteen emerging market economies affirmed that the exchange rate changes affect most of the monetary policy reaction functions. Sek (2009) on studies in three East-Asian countries, confirmed tiny response in the monetary policy functions pertaining to exchange rate changes. Inoue and Hamori (2009) affirmed that monetary policy reaction of India is influenced by exchange rate changes. In view of ASEAN-5 countries, Gan and Kwek (2010) evidently defined that exchange rate changes have significant impact on the policy reaction functions. Likewise, in Thailand, (Luengwilai, 2012) detected the policy reaction function was reactive to exchange rate changes. Meanwhile, in Indonesia, Adenan (2014) identified that the policy reaction function is influenced by exchange rate changes. By the same token, Khalid et al. (2014) validated the importance of exchange rate changes in the policy reaction function of Malaysia.

In particular, focusing on ASEAN-5, (Manogaran and Sek, 2016) applied two different methods, the nonlinear autoregressive distributed lag (NARDL) using yearly time-series data coupled with a pooled mean group (PMG) using yearly panel data. On another study, (Manogaran and Sek, 2017) employed the NARDL method to estimate the quarterly time-series data for individual ASEAN-5 countries. To sum up, the outcome of both these studies revealed the influences of exchange rate changes on the monetary policy function. Another key point of these two studies is the presence of asymmetric effect of exchange rate changes on the policy reaction function.

3. Data, Model and Method

In this study, we apply the threshold regression (with fixed-effects) model to estimate the panel data consist of five countries namely Indonesia, Malaysia, Philippines, Thailand and Singapore which is so-called the ASEAN-5. The panel threshold regression method is utilised to investigate how influential exchange rate (external shock) can affect the monetary policy reaction function in pre-(1980Q1-1996Q4) and post-(1999Q1-2015Q4) crisis periods. As our intention is to compare monetary policy between two policy regimes (pre- and post-crisis). Thus, the financial crisis term (1997-98) is excluded in our analysis as to avoid any biasness in the results.

This method is designed for a balanced panel, hence we use 320 and 335 quarterly balanced datasets extracted from Datastream for the pre- and post-crisis analysis respectively. The main variables are namely, central bank interest rate (INTR) in percentage, consumer price index (CPI) in index form, gross domestic product (GDP) in US dollar and real exchange rate (REER) per US dollar. All the variables involved are converted to natural logarithm, except the interest rates which are in percentages form with the purpose to streamline the data. Further, CPI inflation (CINF) is obtained using log CPI (LCPI) deviates from its lagged 4 (proxy for annual rate). On the other hand,
output gap (GAP) is constructed as log gross domestic product (LGDP) deviates from its trend obtained via the Hodrick-Prescott filter (Hodrick and Prescott, 1997), which typically used by many researchers in generating a smooth estimate of the long-term trend component in GDP series and recognised as potential GDP. Then, the exchange rate changes (DLREER) is the first-differenced of log REER.

In order to consider threshold effect of exchange rate changes (exogenous variable) relative to inflation variation and output gap (regressors) on monetary policy reaction function (dependent variable), we estimate the policy reaction function by utilising the panel static threshold model developed by Hansen (1999). In such the model takes the following form for a single threshold model (equation (2)) and double threshold model (equation (3)): 

\[
\begin{align*}
\text{INTR}_t &= \alpha_i + (\beta_1 \text{CINF}_t + \lambda_1 \text{GDP}_t)(\text{DLREER}_t \leq \gamma) + (\beta_2 \text{CINF}_t + \lambda_2 \text{GDP}_t)(\text{DLREER}_t > \gamma) + \epsilon_{it} \\
\text{INTR}_t &= \alpha_i + (\beta_1 \text{CINF}_t + \lambda_1 \text{GDP}_t)(\text{DLREER}_t \leq \gamma_1) + (\beta_2 \text{CINF}_t + \lambda_2 \text{GDP}_t)(\text{DLREER}_t > \gamma_1) + (\beta_3 \text{CINF}_t + \lambda_3 \text{GDP}_t)(\text{DLREER}_t \leq \gamma_2) + \epsilon_{it}
\end{align*}
\]

Where, \( \beta \) and \( \lambda \) indicate the regression slope of CINF and GDP respectively, \( \gamma \) indicates the threshold values, \( \epsilon_{it} \) indicates the error term, \( i : 1 \leq i \leq n \) indicates the cross-section and \( t : 1 \leq t \leq T \) indicates the time. The assumption holds that the threshold variable (DLREER) is not time-invariant, error term \( \epsilon_{it} \) to be independent and identically distributed (iid) with zero mean and finite variance. In equation (1) the observations are divided into two regimes depending on whether the threshold variable (DLREER) is smaller or larger than the threshold value, \( \gamma \). While in equation (2) the thresholds are ordered so that \( \gamma_1 < \gamma_2 \) and there are either no thresholds, one threshold, or two thresholds.

Prior to the estimations, panel unit-root tests are performed to check the stationarity of each variable. Then, to decide a well-suited model, the significance of threshold effect is observed with a null hypothesis of \( H_0 : \beta_1 = \beta_2 \) and \( H_0 : \lambda_1 = \lambda_2 \) referring equation (2) while \( H_0 : \beta_1 = \beta_2 = \beta_3 \) and \( H_0 : \lambda_1 = \lambda_2 = \lambda_3 \) based on equation (3). In such the rejection of the first or the second null hypothesis implies the existence of single threshold effect or double threshold effect respectively.

4. Results and Findings

Firstly, a battery of panel unit-root tests are employed namely the Levin, Lin & Chu test (Levin et al., 2002), (Breitung, 1999), Im, Pesaran and Shin test (Im et al., 2003) and Fisher-type test using ADF and PP (Choi, 2001); (Hadri, 2000); (Maddala and Wu, 1999). All these tests are with individual intercept and trend specification as well as automatic lag length selection using Schwarz Info Criterion (SIC) to measure the unit-root properties of the datasets. The results of unit-root tests for both sub-periods and whole sample period, rectify all variables (INTR, CINF, GDP, DLREER) are stationary at levels or I(0). As to limit space our results are omitted here.

Thence, we move on using these stationary variables to estimate the monetary policy function up to triple threshold model with specification using a bootstrap of 300, grid of 400 and trim of 0.1. This bootstrap procedure is simple to compute and beneficial in assessing the variability of the estimated p-value in order to construct the related intervals with the presence of boundary issues. Hence, by comparing the three breaks in each pre- and post-crisis periods (refer Table 1) suggesting a different number of threshold for different sub-periods (refer Table 1). Where in the pre-crisis, the single threshold effect’s p-value (0.8600) for F-statistics (0.48) is insignificant and \( H_0 \) cannot be rejected. Consequently, the double threshold effect’s p-value (0.0033) for F-statistics (12.27) showing a strong rejection of \( H_0 \). This affirming the double threshold model of estimation is appropriate. Turning to the post-crisis, the results of p-value (0.0000) for F-statistics (19.34) is strongly significant, hence the single threshold model is adequate.

| Test | Pre-crisis | Post-crisis |
|------|-----------|------------|
| Single Threshold F-stat | 0.48 | 19.34*** |
| Bootstrap p-value (10%, 5%, 1% critical values) | 0.8600 | 0.0000 | (6.28, 7.72, 14.96) | (5.25, 6.69, 8.64) |
| Double Threshold F-stat | 12.27*** | 1.99 |
| Bootstrap p-value (10%, 5%, 1% critical values) | 0.0033 | 0.1633 | (6.03, 7.06, 9.49) | (2.26, 3.18, 6.54) |
| Triple Threshold F-stat | 1.65 | 2.61 |
| Bootstrap p-value (10%, 5%, 1% critical values) | 0.9300 | 0.5667 | (10.63, 13.21, 16.26) | (6.36, 8.19, 10.33) |

Note: *** denotes the statistical significance level at 1%.
Under the above circumstances, Table 2 reports the results of estimations for pre- and post-crisis with a double threshold and single threshold respectively. The varying values of coefficients in each threshold breaks are validating the asymmetric effect of changes in the exchange rate (DLREER) on monetary policy reaction function (INTR) in both sub-periods.

By observing the results of pre-crisis, the double threshold is showing the influences of DLREER on INTR in three threshold breaks (DLREER ≤ -0.0070, -0.0070 < DLREER ≤ 0.0126 and DLREER > 0.0146). However, the influences of changes in exchange rate towards monetary policy reaction function can only be noticed when the threshold value is -0.0070 < DLREER ≤ 0.0126, where the coefficient of DLREER is with a larger value and stronger adverse effect on INTR. This means appreciation of ASEAN-5 currencies leads the policymakers to reduce the policy rate at this threshold value. Apparently, this situation arises because most of these countries were implementing a fixed exchange rate system in the pre-crisis period. So, under a more rigid exchange rate system, the monetary policy is only allowed to react in a narrow interval when the exchange rate appreciation is between -0.7% and 1.26%.

On the other hand, in the post-crisis, the single threshold model illustrates the threshold in two different breaks namely DLREER < -0.0231 and DLREER > -0.0231. In this sub-period, the exchange rate changes trigger strong significant reaction in monetary policy function in all threshold breaks but with different magnitudes. This is a situation when the DLREER threshold value is on the lower side (DLREER ≤ -0.0231) the policymakers are responding towards appreciation of local currency through the implementation of a loosening monetary policy system by reducing the policy rate in order to boost the economic growth. In comparison, when the DLREER threshold value is on the higher side (DLREER > -0.0231) or depreciation rate is larger than 2.31%, the policymakers tend to react towards depreciation of local currency by increasing the policy rate through implementation of a tightening monetary policy as a measure to control the increase in price level or inflation. The inconsistent movements in the policy function based on exchange rate changes at different thresholds clearly demonstrating policymakers’ actions to limit the exchange rate changes. Although, aftermath financial crisis, the majority of the ASEAN-5 countries adopted inflation targeting policy accompanied by free-floating exchange rate system. Yet, our results evidently proving the intervention of policymakers to restrain the fully-flexible exchange rate system. Hence, confirming the existence of ‘fear of floating’ behaviour in ASEAN-5.

| Pre-crisis | Coefficient | Post-crisis | Coefficient |
|-----------|-------------|-------------|-------------|
| DLREER ≤ -0.0070 | 50.3555*** | CINF | 35.4459*** |
| -0.0070 < DLREER ≤ 0.0126 | 0.7584 | GAP | -2.2546 |
| DLREER > 0.0146 | 0.0146 | Threshold: DLREER ≤ -0.0231 | 0.0140, 0.0149 |
| 95% confidence interval | (0.0140, 0.0149) | 95% confidence interval | (-0.0231, -0.0230) |
| C | -5.9556 | Threshold: DLREER ≤ -0.0231 | -215.9636*** |
| -12.2960 | DLREER > -0.0231 | 7.5044*** |
| CINF | -0.0231 | No. of observations | 320 |
| No. of observations | 35 | No. of groups | 5 |
| No. of groups | 5 |

Note: *** denotes the statistical significance level at 1%.

Besides that, we can ascertain that there is a strong direct response from inflation variation (CINF) towards policy function (INTR) in both sub-periods when exchange rate changes being the threshold variable. Such strong reaction of policy function on inflation variation proving that achieving low inflation rate is the main policy target of ASEAN-5. Where the policymakers tend to increase the policy rate in response to close the positive inflation gap with respect to achieving the targeted low inflation rate. On the contrary, there is no effect of the output gap (GAP) towards policy function (INTR) in any of the regimes relative to exchange rate changes. At these threshold breaks, the output gap is not a matter of the policymakers, probably because the economy of ASEAN-5 has achieved sustainable output growth as per targeted level throughout the sample periods of our analysis.

5. Conclusion

By using the panel threshold approach suggested by Hansen (1999) in either pre-and post-crisis case, we explored the existence of asymmetric effects of exchange rate changes (exogenous variable) on monetary policy reaction function of ASEAN-5.

In such, before the crisis, a double threshold effect captures only one limited threshold level of exchange rate changes that trigger significant reaction in the policy function. Alternatively, aftermath crisis, the single threshold effect illustrating all threshold level of exchange rate changes actively responding in diversified magnitudes to the policy function. Hence, we may conclude that external shock (exchange rate changes) plays a very crucial role in the monetary policy reaction function of ASEAN-5 especially aftermath crisis. Where most of these countries adopted...
inflation targeting framework co-exist with flexible exchange rate system. Despite, declaring themselves as free-floaters, the monetary policy reaction function is still behaving asymmetrically towards exchange rate changes. This implies that the policymakers are accommodating the policy function to stabilise the exchange rate changes which obviously remarking the presence of ‘fear of floating’ behaviour among the policymakers of ASEAN-5.

In the final analysis, comparing the threshold effect of exchange rate changes relative to inflation variation and the output gap in triggering the monetary policy reaction function. In both policy regimes, the inflation variation has stronger positive influences on monetary policy function, whereas output gap is negligible. Therefore, we may summarise that the utmost objective of policymakers in ASEAN-5 is to achieve price stability through inflation targeting policy aftermath crisis.

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