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*Genome Medicine* 2013, 5:69  doi:10.1186/gm473

Verena Heinrich (verena.heinrich@charite.de)
Tom Kamphans (Tom@kamphans.de)
Jens Stange (stange@math.hu-berlin.de)
Dmitri Parkhomchuk (parkhomc@gmail.com)
Thorsten Dickhaus (dickhaus@math.hu-berlin.de)
Jochen Hecht (hecht@molgen.mpg.de)
Peter N Robinson (peter.robinson@charite.de)
Peter M Krawitz (peter.krawitz@gmail.com)

ISSN 1756-994X

Article type Method

Submission date 6 February 2013

Acceptance date 22 July 2013

Publication date 31 July 2013

Article URL http://genomemedicine.com/content/5/7/69

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Estimating Exome Genotyping Accuracy by Comparing to Data from Large Scale Sequencing Projects

Verena Heinrich¹, Tom Kamphans², Jens Stange³, Dmitri Parkhomchuk¹, Jochen Hecht⁴,⁵, Thorsten Dickhaus³, Peter N Robinson¹, Peter M Krawitz¹,⁵*

1) Institute for Medical Genetics and Human Genetics, Charité Universitätsmedizin Berlin, Augustenburger Platz 1, 13353 Berlin, Germany
2) Smart Algos, Retzbacher Weg 83, 13189 Berlin, Germany
3) Department of Mathematics, Humboldt-University Berlin, Unter den Linden 6, 10099 Berlin, Germany
4) Berlin-Brandenburg Center for Regenerative Therapies (BCRT), Charité Universitätsmedizin Berlin, Augustenburger Platz 1, 13353 Berlin, Germany
5) Max Planck Institute for Molecular Genetics, Ihnestraße 63-73, 14195 Berlin, Germany

*Corresponding author
Abstract

With exome sequencing becoming a tool for mutation detection in routine diagnostics there is increasing need for platform-independent methods of quality control. We present a genotype-weighted metric that allows comparing all the variant calls of an exome to a high-quality reference dataset of an ethnically matched population. The exome-wide genotyping accuracy is estimated from the distance to this reference set, and does not require any further knowledge about data generation or the bioinformatics involved. The distances of our metric are visualized by non-metric multidimensional scaling and serve as a standardizable score for the quality assessment of exome data.

Background

In recent years, next-generation sequencing (NGS) based exome screens have become an invaluable tool in Mendelian disease gene discovery and are now being introduced as clinical diagnostic tools for genetic disorders of high phenotypic and genetic heterogeneity [1, 2]. Various solutions for exome enrichment and sequencing exist and numerous algorithms for sequence read mapping and variant detection are in use [3-12]. There are recommendations for sequencing depth and benchmarks for the distribution of sequence read coverage over the target region. The common core of the diverse approaches to sequence exomes represent the consensus coding sequences as defined by the CCDS project [13]. The majority of publications [12], which are related to this field, seem to confirm that a mean sequencing depth of this target region with high quality short sequence reads should be above 50 fold and more than 90% of the CCDS exons should be covered by at least ten sequence reads for diagnostic purposes. Another recently introduced parameter for the quality assessment of multiple read alignments is the variance of the ratio of reads that support the alternate allele at
heterozygous positions [14]. The lower this variance, the lower the error rate to be expected from amplification artifacts. The ratio of transitions versus transversions (ti/tv) and the proportion of variants that are already listed in databases of genetic variation such as dbSNP [15], are measures of quality that may be applied to the entire variant call set of an exome. The ti/tv ratio should be close to 3:1 for the CCDS exons and the proportion of singletons, should be below 10% [16]. However, the ti/tv ratio is influenced by the target region, whereas the number of novel variants may depend on the background population. The higher the amount of non-coding variants, the lower is the ti/tv ratio and higher proportions of novelvariants may be observed if the sequenced individual is of a population that is poorly represented in the variant databases.

Although these parameters may serve as valuable indicators for quality they do not directly indicate the accuracy of a sequenced exome and to our knowledge there is no criteria for assessing whether the variants identified by whole exome sequencing represent a comprehensive list. Specifically, it is not possible to estimate an exome-wide false positive or negative rate for variant detection that is purely based on the quality scores of genotype calls. Sequencing technology-specific error signatures [17] can yield artificial variant calls of erroneous high quality and result in an underestimated proportion of false calls, while poorly adjusted bioinformatics pipelines for data processing may lead to missed calls. In most NGS studies a Phred-like quality score is provided for each called genotype. This score describes the confidence in a genotype call. Based on a certain likelihood model for genotypes, the Phred score represents the probability that the genotype call is wrong, given the reads in an alignment (Phred score=$-10\log_{10}Pr(wrong\ genotype)$). In a model that assumes e.g. a Bernoulli distribution of the sequence reads at a heterozygous position, the Phred score of a heterozygous genotype would decrease the more the ratio of reads supporting the alternate allele deviates from the expected value of 0.5. However, this quality score not only depends on the raw data but also on the mapping algorithms and probability models that were used for
variant calling. That means processing the same raw data by different bioinformatics pipelines may result in different distributions of quality scores suggesting different genotyping error profiles for the same exome. Even variant calling approaches that are based on similar Bayesian methods do not yield the same genotype probabilities due to different priors [18] and methods of quality score recalibration cannot completely adjust for that effect (Table 1).

In order to enable interoperability and platform independence, we have developed a method to measure the accuracy of a set of variants by assessing its composition. In our metric the distance between sets of variants from two exome samples is computed without considering genotyping quality scores. The basic idea is that the distance between variant sets of comparable quality is closer than the distance between variant sets of very different quality. In our comparison the variant data of individuals of the 1000 genomes project [19] serve as a gold standard and we refer to them as the reference set. If the genotype concordance between the reference variant call set and the test variant call set is high, this points to a comparable sequencing quality. We will show in the following that the distance of a test sample to this high quality reference data is an indicator for the genotyping accuracy of the exome.

**Materials and methods**

**Generation of exome data, accession of reference data and data processing**

Genomic DNA of European individuals was enriched for the target region of all human consensus coding sequence (CCDS) exons with Agilent's SureSelect Human All Exon Kit according to the manufacturer's protocol and sequenced on an Illumina HiSeq 2000 yielding more than 5 Giga bases raw sequence data per exome. The Charité University Medicine ethics board approved this study that is conform to the Helsinki Declaration and we obtained informed consent of all participants.
Publicly available NGS raw data and variant calls of n = 1063 individuals of different populations were downloaded from the ftp server of the 1000 genomes project [19]. Exome variants of these individuals served as reference variant sets in our work. Exome variants of the 5000 exomes project and of de Ligt et al. were used for testing the accuracy predictions [20, 21].

Exomes of test samples were enriched with Human All Exon SureSelect baits from Agilent and sequenced on Illumina Genome Analyzer IIx and Illumina HiSeq 2000 as 100bp single-end reads or paired-end reads according to the manufacturers’ protocols. Short sequence reads were mapped by novoalign or bwa to the reference sequence GRCh37. Variants were detected with default settings with SAMtools [22] or GATK [10] on bam-formatted alignments [23]. Variant Calls in VCF format [24] were restricted to single nucleotide changes and to the consensus exome target region of the 1000 genomes project. Additionally, sites, which were classified as technical artifacts by the 1000 genomes project, were ignored.

**Distance function**

The distance $d_{ij}$ between any two samples $x_i$ and $x_j$ for all positions $k$ in the target region (exome), where the called genotypes differ from the reference sequence in at least one sample, can be calculated by a weighted indicator function $I \left( x_i(k), x_j(k) \right) \ast W_{ij}(k)$, with

$$l_{ij}(k) = I \left( x_i(k), x_j(k) \right) = \begin{cases} 1, & \text{if } x_i(k) = x_j(k) \\ 0, & \text{if } x_i(k) \neq x_j(k) \end{cases}$$

and $W_{ij}(k) = \frac{2}{f(x_i(k)) + f(x_j(k))}$.

This means that for the same genotypes the Indicator is weighted by the reciprocal of the genotype frequency $f(x_i(k))$ that is based on the reference set with an appropriate background population. To give an example, a genotype for individual $i$ at a given position $k$,
$x_i(k = \text{chr6: 79595096}) = C/C$, would refer to a genotype frequency $f(x_i(k)) = 0.999$, if 1 out of 1000 individuals in the reference set differ from this genotype.

For genotypes that were present only in the test sample, but not observed at all in the reference set, we set their frequency to $1/\text{total number of individuals in the reference set}$. Based on that the distance $d_{ij}$ is defined as

$$d_{ij} = 1 - \frac{1}{C_{ij}} \sum_k I_{ij}(k) * W_{ij}(k)$$

where $C_{ij} = \sum_k W_{ij}(k)$ is used as a normalizing constant.

Therefore, a disagreement at a position of low variability in the reference set contributes more to the total distance than one at a highly variable position.

In the resulting distance matrix $D$, pairs of individuals who are ‘closely related’ can be distinguished from those who are distinctly apart by lower distance values. Thus, a distance $d_{ij} = 0$ means total agreement of all genotypes and a distance value of $d_{ij} = 1$ means total disagreement of all genotypes.

**Visualization of distance matrices by non-metric multidimensional scaling**

The output of the above-described pairwise comparison of variant sets is a high-dimensional distance matrix with given distances or dissimilarities between pairs of individuals, which satisfy all conditions of a metric. To represent the dissimilarities as distances between points in a low-dimensional space, we used a statistical technique named non-metric multidimensional scaling (MDS), that is a visualization methods such as principal component analysis (PCA) or metric MDS. However, in contrast to PCA and metric MDS, non-metric MDS does not make any assumptions about the distribution of the underlying high-dimensional data. With a pre-specified number of dimensions for the embedding $\varphi$ and an appropriate initial configuration, the $isoMDS$ function of the \texttt{MASS} R-package was used to
minimize the goodness of fit, called stress $S$, of Kruskal and Shepard[25]. To promote readability and an easy interpretation of the data, we choose a standard 2-dimensional embedding with

$$S(x_1, \ldots, x_n, \varphi) = \sqrt{\frac{\sum_{i=1,sj}^n (d_{ij} - \| \varphi(x_i) - \varphi(x_j) \|)^2}{\sum_{i=1,sj}^n \| \varphi(x_i) - \varphi(x_j) \|^2}}$$

where $\| \cdot \|$ defines the Euclidean norm.

**Down sampling of raw data and simulation of genotyping accuracy**

For coverage-adjusted comparisons we randomly removed sequence reads from the original alignments. Variants were recalled on these down-sampled exomes as described above. By genotyping accuracy we describe the percentage of the entire exome that was correctly genotyped, being the sum of true positive calls and true negative calls divided by the entire size of the target region. For our simulations we assume that the reference set has a genotyping accuracy of 100% and introduce genotyping errors at random positions. As most of these positions have low variability in the reference set, the contribution of genotyping errors to the distance function can be approximated by adding twice a binomial distributed random variable, $X \sim B(N, p) \ast 2$, with probability $p$ equals the specified genotyping error and the number of trials $N = 2.8 \ast 10^7$ bp is the total size of the exome, to the normalizing constant $C_{ij}$.

**Computation of the standardized dissimilarity score and reference curve**

Distances between all individuals of the reference set were measured and the averaged values of the median and interquartile range (IQR) of all columns of the distance matrix were
computed to standardize the median of a test sample. The median of the distances from a test sample to all individuals of the reference set is computed and subtracted by the precalculated median of the reference set and divided by the IQR of the reference set. The reference curve and both, 5% and 95% -quartiles, for the standardized dissimilarity score (SDS) were computed for the reference set and simulated data sets of decreasing error groups.

Results

An error sensitive genotype-weighted metric

Like any metric, the distance measure that we use to compare different sequences of a set of test samples induces a topology. Variant calls, which describe the measurable differences between samples, represent true genetic variation, as well as genotyping errors. The subject of our work is the quality assessment of a set of exome genotypes, thus our metric needs to induce a topology that is sensitive to sequencing errors while being robust to true genetic differences. By using a weighting method for genotypes that uses their frequency of occurrence, we achieved higher precision in accuracy prediction than compared to an unweighted hamming distance (Additional file, Figure S1). If two samples are not the same at an exonic position, which is highly constrained in the population, this contributes more to the total distance, as such an event has a higher probability of being a genotyping error than divergent genotypes at a highly variable site of the exome. When we compare two exomes our metric works on all genotypes that have been called in these samples. The genotypes are weighted by the degree to which the genotype is constrained in the population. Though several definitions for measuring the degree of genotype conservation have been suggested [26, 27], most variable positions in the human genome are biallelic and for simplicity we approximate the conservation of a genotype by the inverse of its frequency. Thus the differences in two variant sets are weighted by their respective genotype frequencies and the detection of many rare variants in a test sample therefore points to a higher proportion of false
positive genotype calls. On the other hand, if many variants that are common in the population are not detected in a test sample, this points to a high false negative error rate. By this means we compute a matrix that contains the distances of the test sample to all the samples of a reference set as well as their mutual distances. These distances are a result of a function that works on the entire exomic target region, as defined by the 1000 genomes project and may therefore be viewed without distortion only in the multivariate, exomic space. We have implemented and tested our method using whole exome data, but it could be applied to other types of high-throughput sequencing data. However, the precision of predicting the accuracy of genotyping decreases for smaller target regions (Supplemental Figure 2).

**Non-metric MDS is best suited for distance visualization**

As the distance is based on multiple variables of weighted categorical data, the visualization in a plane requires a transformation. We tested several standard techniques of data visualization and found that non-metric multidimensional scaling, MDS [28], showed the best characteristics in conveying the differences in genotyping accuracy in two dimensions (Additional file, Figure S3). We therefore project the exomic distance matrix into two artificial dimensions of Φ1 and Φ2 that have the smallest loss of information [29-31].

The reference samples of the 1000 genomes project form clusters according to their ethnicity (Additional file, Figure S4) and for any test sample we chose the closest cluster as the best matching reference set. Samples from the same population background form homogeneous clusters in non-metric MDS scaling indicating a comparable genotyping quality (Figure 1 A). We then analyzed two test samples of European descent but of unknown genotyping accuracy and included them into the MDS projection of all CEU individuals from the 1000 genomes project, which is shown in Figure 1B. Except for one representative recalled sample, NA06986, all individuals of the CEU reference set are now displayed as black circles,
whereas the two exome test samples are represented by the colored triangles. Although the mean sequence coverage of the exome target region is above 60 x for both test samples, they clearly differ in their mean distance to samples from the reference set: The second sample is close to the cluster formed by the individuals of the reference set, while the first sample is an outlier indicating inferior quality. This considerable difference in the mean distance is remarkable given the high sequencing depth and a comparable ti/tv ratio of 3.2 (Table 1). Also the proportion of variants found in dbSNP is around 97%, which is comparable to NA06986. Only the variance of the heterozygous allele frequencies, that increases with a growing number of artifacts from the amplification steps during the library preparation, suggests a lower quality for sample 1 with var(het AF)=0.017 compared to 0.012 in test sample 2 and 0.013 in NA06986 at the same mean coverage [14].

**Visualization of exomes of different genotyping accuracy**

We measured the mean distance to the reference set for 85 exomes of European descent that were all analyzed in the context of NGS research projects. The two samples displayed in Figure 1B illustrate the extreme spectrum of the mean distance to the reference set that we encountered. We hypothesized that variants often detected in exomes of the 1000 genomes project, but not in the outliers of our test samples, might point to a subset that requires high data quality to be properly detected and that might explain partly the separation from the reference cluster. Figure 2A displays the value pairs of variant allele frequencies that are based on 85 CEU individuals from the 1000 genomes project and 85 exome test samples. For technical replicates one would expect all allele frequency value pairs to lie close to the diagonal in this kind of visualization. For two sample sets of equal size that are drawn from the same population one would expect a certain degree of variance in the measured allele frequency that is simply based on the finite sample size: Given the allele frequencies one would expect for a sample size of 85 that about 95% of the frequency value pairs fall inside
the boundaries of the displayed ellipse based on a Bernoulli distribution. However, in our case there are considerably more outliers than expected by chance.

**Characteristics of sequence variants with high error rates**

We looked for similarities of these outliers and computed the GC content of 100 bases flanking the variant alleles that were present in more than half of the individuals analyzed in the 1000 genomes project but in only one or less of our analyzed samples. The distribution of the GC content of these variants clearly deviates from the distribution that one would expect for randomly located variants in the exome (Figure 2B).

To investigate the reasons of the higher false negative error rate for variants in a GC-rich sequence context we computed the mean read coverage of the target region depending on the GC-content. Figure 2C shows that the distribution of the coverage is smaller for the test sample 1 compared to the test sample 2 and NA06986 that was sampled down to a comparable mean coverage. Thus, the higher distance of test sample 1 to the reference set is partly due to a critically low sequence read coverage of regions with an extreme GC-content. Benjamini et al. studied the bias caused by GC content depending on read coverage in detail for Illumina sequencing data and showed that it also varies between technical replicates [32]. This also means that two samples may have different genotyping accuracy for the whole exome although they have been processed by the same protocol.

**The distance to the reference set versus coverage and error rates**

We hypothesized that the distance to the high quality reference set should grow when the amount of raw sequence data decreases. We therefore successively reduced the sequence coverage in the raw exome data of NA06986, called variants anew and observed an increasing distance to the reference set (green to blue circle in Figure 3). A decreasing sequencing coverage will not only yield an increasing rate of false negative genotypes but will also
increase the rate of false positive calls. It is more likely that a sequencing error will be called as a heterozygous variant, and a heterozygous variant as homozygous if the position is only covered by a few reads. We then analyzed how an increasing false positive error affects the distance to the reference set by simulating detection artifacts that were randomly distributed over the target region and added to a sample from the reference set. The triangles in Figure 3 show that the data points follow a trajectory that departs from the reference cluster with growing error rate. It has to be noted that the visualization of multiple versions of the same test sample that differ only in coverage or error rates in a single MDS plot contorts the relative distances. All the depicted simulated data points in Figure 3 originate from the same data set and are therefore self-similar. The self-similarity is high for a low error rate and high coverage and decreases with increasing error rate and decreasing coverage, as the divergent trajectories of circles and triangles indicate. We also obtained similar results in the analysis of simulated data sets of other ethnic backgrounds (Additional file, Figure S5).

A genotyping error of 0.00001 corresponds to an expectation of one genotyping error in approximately 100kb of the target region. Two randomly chosen samples from the reference set would differ in about 100 positions in such a window of 100kb [21]. The samples with the simulated error rate begin to separate from the high quality cluster for error rates above 0.00001, which corresponds to a positive predictive value of 0.99 (number of true positive divided by number of positive calls). Interestingly, the positive predictive value that was reported by Tennessen et al. for the variant calls of the 5000 exomes project is between 0.97 and 0.98 [21]. The resolution of our visualization techniques is therefore sufficient to display these qualitative differences.

Comparison of exome data from different NGS studies

In contrast to the 1000 genomes project, the genotype calls from the 5000 exomes project were publicly available only in a collapsed form as genotype frequencies for European
Americans and African Americans and not as separate variant sets for each sample. In addition to our in-house exome data we also analyzed the distances to the reference cluster for exomes that we simulated based on the genotype frequencies from these European Americans and 100 exomes that were already studied by de Ligt et al [20]. Figure 4A shows the distribution the standardized dissimilarity score, SDS, which represents a normalized distance of a test sample to the reference cluster. The mean SDS for the simulated exomes of the 5000 exomes project is comparable to our exome data and lower than the SDS of the de Ligt exomes. The smaller variance of the SDS distribution in the 5000 exomes samples that points to a higher self-similarity is due to the simulation process that did not properly represent the haplotype substructure of the data. The higher mean SDS of the de Ligt et al. data is mainly explained by out-dated genotyping algorithms with higher error rates and a lower mean coverage of the target region. Figure 4B depicts the coverage distribution over the exome and additional quality parameters for an exemplary sample from de Ligt et al. and our two test samples.

By simulating increasing sequencing noise for all exome data sets of the 1000 genomes, we derived distributions for the mean distances of the original data. Based on these distributions we computed a reference curve for the SDS of an unknown sample that correlates with its exome-wide accuracy (Figure 5). The SDS that is measured for our test samples may be used for estimating their genotyping accuracy by intersecting with the reference curve. For test sample 1 the mean distance to the reference set is 0.29, which corresponds to an estimated genotyping error of 0.0001. In contrast the SDS of 0.11 for test sample 2 indicates an error rate that is much closer to that of the 1000 genomes project. Interestingly the distance to the reference set shows characteristics of a phase transition, when the contribution of genotyping errors exceeds the genetic variability between individuals. We also checked the validity of our approach by deriving the genotyping accuracy by a complementary method. In Heinrich et al. we analyzed the effect of amplification steps during sample preparation and derived rates of
genotyping errors from technical replication [14]. The SDSs for these replicates indicate genotyping accuracy for the exomes between 99.99% and 99.999%, which is in good agreement with the previously computed accuracy.

Thus, the SDS is a parameter that is derived from the composition of a variant set and that is even more powerful in predicting the data quality than other QC parameters such as coverage distributions, which require access to the read alignments (Table 1).

We tested the influence of the sequencing platform on the error prediction by our approach by restricting the reference set to samples that were sequenced with the same technology (Illumina). Although the visualization in metric MDS clearly shows that the test samples are closest to the Illumina samples from the reference set (Additional file, Figure S3), the effect of the sequencing platform on the accuracy prediction is marginal (Additional file, Figure S6).

The SDS is therefore robust and can be applied to predict the quality of genotyping data from different sequencing technologies. We used the platform independence of our approach to analyze the quality of a sample from the 1000 genomes project, NA12878, that was recently resequenced by a new technology (Proton, Ion Torrent, aligned with TMAP and genotyped with variantCaller). In Figure 6 the distance of this Proton variant set of NA12878 to the reference set is compared to a variant set of the same individual that is based on Illumina raw data that was down-sampled to a mean coverage of 30 fold as described above. This figure visualizes how the quality of two exomes of the same individual that were generated on two different sequencing platforms and processed by different bioinformatics pipelines may be interpreted at a glance. We used the SDS to estimate the genotyping accuracy and predicted 99.9% for the low coverage Illumina exome and above 99.99% for the Proton exome, which is in good agreement with the values that are based on Sanger validation for this sample.
Discussion

In this work we described a new approach to assess the accuracy of variant calls from NGS studies. The genotyping accuracy for variant calls, that is genotypes which differ from the true sample sequence, has been estimated in large scale NGS-based projects such as the 1000 genomes project [19] or the exome sequencing project [21] and comparisons of NGS platforms [33]. In these projects samples were sequenced to a very high mean coverage on different sequencing platforms and the reported variants represent an intersection of technical replicates and independent analysis pipelines. Even in these high quality data sets, up to about 2% of the variants cannot be validated if re-sequenced by a complementary approach such as ABI Sanger. In such a high quality exome one detects around 15,000 single nucleotide variants per 30MB coding sequence and approximately 300 of them are likely false positive calls which corresponds to an error rate of 300/30*10^6=0.00001.

Based on simulated accuracy groups for variant calls we were able to assess the quality of an exome test sample without detailed knowledge about the applied enrichment and sequencing technology or about the bioinformatics pipeline that was used to align the reads and call the genotypes. The SDS is therefore suitable for a comprehensive quality control in all exome based mutation screens and might turn out especially useful as a criterion for data inclusion in studies that combine exome data of different sources, due to its platform independence. Especially the estimated genotyping error might serve as quality criterion before variants detected in an exome are further analyzed: Only if the estimated error is comparable to that of a high quality reference set such as the 1000 genomes project one would proceed with variant analysis. We envision that our approach to estimate the genotyping accuracy of exomes will facilitate the quality assessment of NGS data.

Software that computes the SDS, visualizes the distances to data of the 1000 genomes project, and predicts the genotyping accuracy is available for download and as a web service at GeneTalk [36].
Abbreviations

1KG: 1000 genomes Project; CEU: Central European; JPT: Japanese; MB: Megabase, 10^6 bp;
NGS: next-generation sequencing; SDS: Standardized Dissimilarity Score; WES: whole
exome sequencing; YRI: Yorubian

Competing interests

The authors declare that they have no competing interests.

Authors’ contributions

VH, PNR, and PMK conceived and designed the study. VH developed the original code, TK
and DP implemented modifications for GeneTalk. JH contributed exome data. JS and TD
supervised and contributed to the statistical analysis. VH, PNR, TD, and PMK wrote the
paper.
All authors read and approved the final manuscript.

Web resources

The URLs for data and methods presented herein are as follows:
NHLBI Exome Sequencing Project (ESP) Exome Variant Server,
http://evs.gs.washington.edu/EVS/
GeneTalk, https://gene-talk.de/qc
ftp server of the 1000 genomes project, ftp://ftp.1000genomes.ebi.ac.uk/vol1/ftp/

Acknowledgements

This work was supported by a grant of the Deutsche Forschungsgemeinschaft (DFG KR
3985/1-1) to P.M.K. We thank Christian Gilissen and Joris Veltman for providing solid
exome data and Michal R. Schweiger and Martin Kerick for cancer exome data for analysis.
The authors would like to thank the NHLBI GO Exome Sequencing Project and its ongoing studies which produced and provided exome variant calls for comparison: the Lung GO Sequencing Project (HL-102923), the WHI Sequencing Project (HL-102924), the Broad GO Sequencing Project (HL-102925), the Seattle GO Sequencing Project (HL-102926) and the Heart GO Sequencing Project (HL-103010).

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Figure legends

Figure 1: Similarities of exome samples: A) The similarity between exome samples is measured with a genotype frequency weighted metric. The similarity matrix is visualized by non-metric multidimensional scaling, MDS, in the two-dimensional plane. Variant sets of individuals of European descent that were analyzed by the 1000 genomes project form a homogeneous cluster, irrespective of the applied sequencing technology. B) Two test samples the same ethnicity that we genotyped according to a standard exome protocol are compared to
samples of the reference set. The larger distance of sample 1 to the reference cluster indicates a lower genotyping accuracy, while the sample 2 points to a high quality.

**Figure 2: Analysis low quality samples.** A) Comparison of variant allele frequencies in different sample sets. Value pairs of genotype frequencies were computed for exomes of the reference set (CEUs from the 1000 genomes project) and compared to test samples of the same ethnicity that are low quality. The ellipse indicates twice the standard deviation assuming a binomial model for the allele frequency $p$. Variants in the right lower quadrant were called with a lower probability in our test samples and are characterized by a GC content that deviates from the expected mean. B) GC Content at false negative positions. Variants that are underrepresented in exomes with a large distance to the high quality reference set are overrepresented in exome regions with high GC-content (violet curve). The green curve indicates the distribution of the GC-content that is expected for an equal number of variants that are randomly drawn from the exome. C) Coverage against GC-content. The mean sequence coverage of the consensus exome varies with the GC-content of the target region. The overall coverage for an exemplary sample from the 1000 genomes project (NA06986) was higher compared to test samples 1 and 2. Test sample 1 has a particularly low coverage in regions with extreme GC-content suggesting a higher error rate.

**Figure 3: Distance to the reference set for decreasing coverage and increasing error rate:** A reduction of the raw sequence amount for a randomly chosen sample of 1KG project, NA06986, reduces the similarity to the reference set (green-blue circles) indicating decreasing genotyping accuracy. A similar effect is observed when simulated genotyping errors are added to the variant calls (red-yellow triangles). A decreasing similarity of the test samples results in divergent trajectories.
Figure 4: Distributions of standardized dissimilarity scores, SDS, for different exomes. A) 100 sample, including test samples 1 and 2, were sequenced on the Illumina platform with mean coverage above 60 x. The mean SDS for the Illumina samples is comparable to the mean SDS of 100 exomes that were simulated based on genotype frequencies from the 5000 exomes project. The variation of these SDS is smaller due to missing haplotype information. The mean SDS of 100 exomes that were sequenced on a Solid platform is considerably higher due to a lower and less uniform sequence distribution over the target region and due to less accurate variant calls (B).

Figure 5: Estimation of genotyping errors from standardized dissimilarity scores. The reference curve with its 5% and 95% quantiles is based on the distances of samples with simulated error rates to the reference set. The standardized dissimilarity score, SDS, of a test sample indicates its error rate by its intersection with the reference curve. The estimated error rate of test sample 1 is considerably higher than of test sample 2 and of NA06986 from the 1KG project.

Figure 6: Comparison of exome data of different NGS platforms. A) Exome genotypes of individual NA12878 were called based on an Illumina read data set of low coverage and a Proton dataset. The Proton exome is closer to the CEU reference data than the low coverage Illumina exome. B) The accuracies of the exome genotype sets were estimated based on the mean distance to the reference set.

Table 1: Comparison of different parameters for quality assessment. Short sequence reads of test sample 1 and 2 and a sample from the 1000 genomes reference set, NA06986, were sampled to comparable mean coverage over the target region. The total number of variant calls decreases with a decreasing coverage indicating an increasing false negative
error rate. Also the mean genotype quality scores decrease with a decreasing coverage indicating an increasing false positive error rate. The ti/tv ratio and the ratio of variants that are present in dbSNP vary only very little with changing coverage. Different priors in the genotyping models of Samtools and GATK result in different mean genotype quality scores for the same alignments. Quality score recalibration with GATK VariantRecalibrator performed on Samtools- and GATK-called variants ads an adjusted quality score, VQSLOD (log odds ratio of being a true variant versus being false under a trained gaussian mixture model). This score is used in GATK ApplyRecalibration to generate a tranche file of highly confidential calls. The percentage of these calls which passed as high-confidential shows an irregular behavior with respect to the mean coverage. The standardized dissimilarity score, SDS, which is computed for the entire variant call set of a sample correlates with its accuracy and allows a sample-to-sample comparison.

Additional file, Figure S1: Distribution of dissimilarities of simulated error groups for different distance metrics. The distance of simulated test samples of different error groups were to the reference set was measured with an unweighted hamming metric and a genotype frequency weighted metric. The variance of the dissimilarities of the test samples is smaller in the genotype frequency weighted metric and thus allows a more precise prediction of the error group.

Additional file, Figure S2: Visualization of distances and estimation of error rates for different target regions. The distances of two test samples (sample 1 of high and sample 2 of medium quality) to the reference data were computed for five different target regions that differ in size. The CCDS exome comprises 29Mb, the human phenotype ontology (HPO) [34] panel contains all exons of genes associated with phenotypic features (5.8Mb), the Kingsmore panel comprising 548 genes of known inherited diseases (1.2Mb) [35], all coding exons of
chromosome 22 (600 kb), and the GPI panel that contains all genes involved in the GPI-anchor synthesis (45kb). The larger the target region, the higher is the number of sequence variants for comparison. This increases the precision of the estimation of the error rates. With decreasing size of the target region the confidence intervals of the reference curve for the standardized dissimilarity score widen. While the different error rates of sample 1 and 2 can be clearly estimated and visualized for the larger target regions, gene panels below 1MB do not allow this assessment any more due to the larger confidence intervals.

Additional file, Figure S3: Data visualization techniques. Comparison of ordination methods for the visualization of the distances of exome genotypes of two test samples and high quality reference samples of a matched background population. The mean distance of test sample 2 with the low genotyping accuracy to the reference samples is larger compared to sample 1 with the high genotyping accuracy for all visualization methods. For PCA and metric MDS a substructure in the reference samples is visible that is specific to the sequencing platform.

Additional file, Figure S4: Exomes of different ethnicities (European CEU, Yorubian YRI, Japanese JPT) form distinct clusters based on their similarity. For a test sample the closest cluster from the 1000 genomes project data is chosen as reference set.

Additional file, Figure S5: The distance of a test sample of the Yorubian reference set increases for a growing simulated error rate.

Additional file, Figure S6: Influence of sequencing platform on error prediction. In contrast to non-metric MDS visualization PCA of the similarities of European samples of the 1000 genomes project reveals some information about the sequencing platform that was used
(A). However, the effect of the sequencing platform for predicting the genotyping accuracy is small. The predicted error rates of the test samples are comparable if the reference set is restricted to specific sequencing platforms (B, C).
Figure 2

- 1000Genomes CEU Individuals
- CEU Test Sample 1
- CEU Test Sample 2
- NA06986
Figure 5

Coverage against GC content

| Sample                  | Mean Coverage |
|-------------------------|---------------|
| CEU Test Sample #1      | 65            |
| CEU Test Sample #2      | 142           |
| NA06986 20%             | 63            |
| NA06986 40%             | 125           |
| NA06986 60%             | 188           |
| NA06986 100%            | 313           |
Non–Metric MDS
With Different Coverages and Error Rates

Mean Coverage of NA06986:
- Mean Coverage: 313
- Mean Coverage: 100
- Mean Coverage: 65
- Mean Coverage: 30
- Mean Coverage: 15

Simulated Error Rates:
- Error Rate: 0
- Error Rate: 0.00001
- Error Rate: 0.0001
- Error Rate: 0.0005
- Error Rate: 0.001

Figure 6
A) Standardized Dissimilarity Score

- 100 Random Samples (5000 Exomes)
- 100 Illumina Samples
- 100 Abi Solid Samples

B) Coverage and Variant Call Statistics

- **Abi Solid Sample:**
  - Mean Genotype Quality: 75.1
  - TiTv Ratio: 1.9
  - Percentage in dbSNP135: 74.8
  - Var(het AF): 0.018

- **Illumina Sample 1:**
  - Mean Genotype Quality: 89.7
  - TiTv Ratio: 3.1
  - Percentage in dbSNP135: 97.1
  - Var(het AF): 0.017

- **Illumina Sample 2:**
  - Mean Genotype Quality: 94.4
  - TiTv Ratio: 3.2
  - Percentage in dbSNP135: 97.4
  - Var(het AF): 0.009
Figure 8
| Sample coverage | Test Sample 1 | Test Sample 2 | NA06986 |
|-----------------|---------------|---------------|---------|
| coverage        | 15 30 50 65   | 15 30 65 100 142 | 15 30 65 100 313 |
| Number of Called Variants: | | | |
| samtools        | 13999 15446 16112 16289 | 15060 16535 17098 17268 17220 | 14071 16426 17683 18078 18332 |
| GATK            | 14597 16065 16813 17083 | 15413 16862 17794 18093 18295 | 13493 16434 18586 19503 21010 |
| GATK recalibrated bam | 14525 16033 16765 17033 | 15348 16823 17787 18095 18292 | 13308 16347 18479 19410 20946 |
| Mean Genotype Quality (GQ): | | | |
| samtools        | 71.8 83.4 88.0 89.7 | 71.3 84.7 91.8 93.4 94.4 | 60.4 77.2 87.3 90.4 94.8 |
| GATK            | 69.8 81.7 87.7 89.8 | 69.0 81.7 91.2 93.5 94.7 | 57.5 73.2 85.1 89.4 94.7 |
| GATK recalibrated bam | 69.6 81.6 87.7 89.8 | 69.0 81.7 91.2 93.4 94.7 | 57.1 73.1 85.1 89.4 94.7 |
| Mean VQSLOD Score: | | | |
| samtools recalibrated vcf | 0.75 14.12 13.69 14.46 | 0.81 14.58 11.32 14.46 14.89 | 13.99 12.69 1.94 15.68 1.86 |
| GATK recalibrated vcf | 11.49 4.21 3.33 3.09 | 9.68 3.33 -0.34 -0.05 0.51 | 1.58 2.15 2.02 1.94 1.59 |
| Percentage of Variants, which were flagged as 'PASS': | | | |
| samtools recalibrated vcf | 81.9 93.8 94.0 93.9 | 80.5 92.3 91.8 92.1 96.5 | 91.6 91.4 85.2 91.2 88.0 |
| GATK recalibrated vcf | 91.8 90.9 91.0 90.9 | 90.4 82.7 81.7 84.4 86.2 | 90.4 85.7 83.3 81.4 76.9 |
| ti/tv-ratio: | | | |
| samtools        | 3.3 3.3 3.3 3.2 | 3.2 3.2 3.2 3.2 | 3.0 3.0 3.0 3.0 | 3.0 3.0 3.0 3.0 |
| GATK            | 3.2 3.2 3.2 3.1 | 3.2 3.1 3.1 3.0 | 3.0 2.8 2.6 2.4 | 2.2 2.2 2.2 2.2 |
| Percentage of Variants present in dbSNP135: | | | |
| samtools        | 97.2 97.2 97.1 97.1 | 97.3 97.4 97.4 97.4 97.4 | 96.1 95.5 95.8 96.1 97.0 |
| GATK            | 97.4 97.0 96.7 96.5 | 97.7 97.4 97.1 96.9 96.7 | 96.2 94.1 91.8 90.4 87.6 |
| Variance of heterozygous allele frequencies: | | | |
| samtools        | 0.021 0.019 0.017 0.017 | 0.020 0.015 0.012 0.010 0.009 | 0.022 0.018 0.013 0.011 0.007 |
| GATK            | 0.022 0.020 0.019 0.018 | 0.020 0.016 0.014 0.013 0.012 | 0.020 0.016 0.012 0.011 0.009 |
| Standardized Dissimilarity Score (SDS): | | | |
| samtools        | 0.73 0.58 0.52 0.52 | 0.58 0.44 0.38 0.38 0.38 | 0.73 0.5 0.38 0.35 0.31 |
| GATK            | 0.56 0.39 0.31 0.29 | 0.42 0.26 0.16 0.13 0.12 | 0.72 0.42 0.25 0.21 0.13 |
| GATK recalibrated bam | 0.57 0.39 0.31 0.29 | 0.43 0.26 0.16 0.13 0.12 | 0.73 0.42 0.26 0.21 0.13 |

Table 1: Comparison of exomes with different coverages.
Additional files provided with this submission:

Additional file 1: SupplementalData.docx, 501K
http://genomemedicine.com/imedia/9159012661041363/supp1.docx