Positive but not negative affect is associated with increased daily drinking likelihood in non-clinical populations: systematic review and meta-analyses

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ABSTRACT

Background: Recent meta-analytical findings indicate that affect regulation plays an important role in alcohol craving, consumption volume, and substance use. However, in view of mixed findings, the affect and drinking likelihood literature remains in need of clarification and consolidation.

Objectives: This systematic review with meta-analyses interrogated the results from peer-reviewed studies among non-clinical populations that examined the relationship between daily affective states and intraday likelihood of alcohol consumption.

Method: A PRISMA guided search of PsychINFO, PsycARTICLES, Science Direct, Wiley Online Library, PubMed, SCOPUS, and JSTOR databases was conducted. Multilevel meta-analyses yielded 11 eligible negative affect studies (2751 participants, 23 effect sizes) and nine studies on positive affect (2244 participants, 14 effect sizes).

Results: The pooled associations between intra-day affect and alcohol consumption likelihood revealed no significant association between negative affective state and drinking likelihood (OR = 0.90, 95% CI [0.73, 1.12]) and that positive affect was associated with increased drinking likelihood (OR = 1.17, 95% CI [1.09, 1.27]). Egger’s test, P-curve, fail-safe N, and selection models analyses suggested that the obtained results were unlikely to be the product of publication bias and p-hacking alone.

Conclusions: Results converge to suggest that, independent of age, affect measure used, and study design, a significant albeit modest relationship between positive affect and alcohol consumption likelihood exists, which does not appear to be the case for negative affect. In conjunction with other recent meta-analyses, current findings help map out a more nuanced understanding of the affect-alcohol/substance use relationship, with potential implications for interventions.

PUBLIC HEALTH SIGNIFICANCE STATEMENT

These comprehensive meta-analyses on the impact of within-day affective states on alcohol consumption suggest that increased positive but not negative affect is associated with greater likelihood of alcohol consumption.

Introduction

It is commonly believed that alcohol has the ability to impact people’s mood (e.g., 1). The cultural association of alcohol with both celebratory (e.g., birthdays) and sad (e.g., funerals) life events in some societies also means that normative beliefs surrounding drinking (2,3) may help sustain the apparent link between affect and alcohol consumption. Empirical evidence, in this vein, suggests that alcohol consumption is associated with increased self-reported pleasurable moods (4–8) and decreased nervousness, stress, anxiety, tension, depression, and self-consciousness (4, 8–12). Accordingly, a sizable body of literature theorizes that, in part, via association, individuals consume alcohol to control, or to moderate, their affective states (13), and that mood management is a commonly anticipated outcome of drinking (14). Drinking to improve positive states, or to cope with negative affect, is also frequently found to be associated with elevated consumption (15–17), and recent meta-analytical findings indicate that emotion regulation abilities and substance use are intertwined (18). At first sight, it may therefore be expected that the association between affect and consumption would be a consistent and fully understood phenomenon. However, to date, research examining the relationship between mood and drinking likelihood has yielded somewhat mixed findings. This contribution therefore aimed to meta-analytically examine the relationship between mood and the likelihood of consuming alcohol on a given day, and to investigate potential moderators of this association.
According to the positive reinforcement theory of alcohol use, people drink alcohol to enhance positive emotions they are currently experiencing (19). Indeed, enhancement drinking motives and positive alcohol expectancies appear to be important drivers of consumption in different age groups (17,20–22). However, findings with regard to the extent to which people’s current positive affective states impact drinking behavior have been varied. While some studies have found that positive mood increases within-day alcohol consumption (23–27), others indicate that positive affect has no impact on daily alcohol consumption volume (28–34) or on alcohol craving (35). In view of this inconsistent support for the positive reinforcement model of alcohol consumption in non-clinical populations, formal interrogation of the findings is required to shed light on the relationship between positive mood and the likelihood of drinking.

In a similar vein, in accordance with the self-medication hypothesis (13) and negative reinforcement models of alcohol use (36), research also documents an association between poor mental health and hazardous drinking (37), such that people may use alcohol as a means of trying to overcome low mood. In line with this assertion, symptoms of depression, anxiety disorder are highly comorbid with alcohol use disorders (38,39). Nevertheless, research in this domain has not produced findings that are as consistent as these established theories would predict. While, in non-clinical populations, there is evidence that negative moods are associated with substance use cravings and increased consumption (40), it has also been found that such affective states do not impact alcohol consumption in people with and without mood disorders (41,42), that there is no association between negative affect and unplanned heavy drinking (43), and that coping with depressive symptoms is unrelated to increased consumption (44). Furthermore, other research indicates that negative affect is inversely related to drinking (45). The existing literature therefore provides mixed support for the self-medication hypothesis in relation to alcohol consumption likelihood; however, to date a comprehensive meta-analytical interrogation of this relationship has not been published.

A potential reason for mixed findings concerning how affect may impact alcohol consumption relates to how outcomes are measured and conceptualized in relevant studies. Alcohol consumption can be assessed methodologically in different ways – e.g., amount consumed, rate of consumption, blood or breath alcohol concentration, alcohol cravings, or drinking likelihood. Previous meta-analytical findings in the field indicate that both negative and positive affect have a tendency to be associated with increased daily alcohol consumption volume (46), that lab-induced negative affect may increase alcohol craving (47), and that enhancement (i.e., to increase positive affect) drinking motives are associated with drinking problems and alcohol use, while coping (i.e., to deal with negative affect) drinking motives are associated with increased alcohol use (48). However, there has been considerably less research regarding the impact of current mood on the likelihood of drinking onset.

Another potential reason for the mixed findings in this research area may relate to the variable use of measures that conceptualize affect in different ways. For example, while some scales conceptualise negative and positive affect as distinct constructs (49,50), others treat affect as a continuous construct that varies in valence and hedonic tone (51,52). Using diverse measures could therefore potentially lead to different interpretations of people’s affective states. As such, when scales such as the PANAS are used, participants may report experiencing both positive and negative emotions at the same time. On the other hand, when scales that conceptualize positive and negative affect as a bipolar continuum are used (51), participants have to choose the affective state which they consider to be stronger at the moment they complete the measure. Yet, there is evidence people can experience both positive and negative affect simultaneously (53) and that difficulties regulating positive and negative affect are distinct (54). It follows that each affect dimension could account for unique (non-shared) variance and, as such, measuring affect as a bipolar continuum could be a potential confound within some previous research. Given the variety of different scales utilized in affect-alcohol research, which may vary in the extent to which they are accurately able to assess people’s current mood, affect scale was examined as a possible moderator in the present study.

Study design may also be a source of variability in previous research and offer insights into possible reasons for inconsistent previous findings. It is possible to examine drinking likelihood by using one-time surveys, daily diary methods, or ecological momentary assessment methods (EMA). However, studies tend to vary in length, with shorter studies potentially missing small effects (55). While some studies have a duration of 7 days (28,56), others can last for up to 30 days (57). Similarly, studies vary with regard to the number of assessments per day, with some using fewer assessments and thereby potentially missing nuances in the ways that daily affect could be associated with consumption patterns. As such, while some studies only require participants to complete one assessment per day (57–59), others ask participants to complete up to 11 assessments (28). Therefore, shorter studies and those using fewer prompts may not be able to capture fully the variability of mood throughout the day
Another important aspect of study design is whether affect was measured prior to consumption. Studies that do not assess (or control for) this may prone to false-positive results (as findings may reflect the impact of alcohol on affective states, rather than the other way round). To examine this possibility, differences in study design need to be considered as part of meta-analyses of the mood alcohol relationship.

Another factor that could be a moderator of the relationship between affect and alcohol consumption likelihood is the age of participants taking part in relevant studies. While most research in the area has been conducted on adults, adolescents may also be prone to drinking in response to particular emotions. For example, studies indicate that adolescents experience both negative and positive emotions more intensely than adults (60,61), and increased affect intensity, in this way, may lead to a greater likelihood of drinking for emotion regulation purposes, compared to adults who may be more likely to deal with their stress and associated emotions in alternative ways (62). Similarly, rates of depression and anxiety (63,64), and response intensity to stress (65) tend to increase during adolescence. Thus, this life stage seems to be a period that is characterized by intense emotional experiences (66) which, taken together with concomitant proneness to risk-taking (67), could potentially contribute to elevated alcohol consumption. Indeed, individuals in their early-to-late teens and early 20s are among the heaviest episodic drinkers (68), although findings on whether increases in negative affect during adolescence could be related to alcohol consumption have been inconsistent (69–74), and recent research pointing to declining alcohol consumption among younger age groups must be acknowledged (75,76). With regards to positive affect, to our knowledge, there is a lack of studies on how this may impact the daily likelihood of alcohol consumption in adolescents, although research that examined this in young adults has yielded varied findings, with studies showing that both high (77) and low (78) levels of positive affect may be associated with alcohol consumption volume. Investigating the extent to which young people’s alcohol consumption likelihood may be associated with particular affective states could therefore yield insights into factors that drive their alcohol consumption.

Finally, wider literature documents the occurrence of a decline effect (79) whereby effect sizes may decrease over time for a variety of possible reasons that include false-positive results, overestimation of effect sizes, under-specification of the conditions of the study, or genuinely decreasing effect sizes (80), as well as methodological and analytical advances (46). As a previous meta-analysis of the relationship between affect and alcohol consumption volume demonstrated a decline effect for studies on positive affect (46), it is possible that studies examining the odds of consumption may also be susceptible to this phenomenon. Therefore, year of publication was also examined as a moderator in this study.

Accordingly, the current review aimed to meta-analytically synthesize empirical findings regarding the intra-day impact of affective states on alcohol consumption likelihood. To inform prevention efforts, this study focused on non-clinical populations as there might be potentially different drivers of alcohol consumption for clinical samples. Since positive and negative affect are distinct but related constructs that both appear to be linked to alcohol consumption (although the direction is unclear due to mixed findings), the decision was made to examine these separately. Several moderators, including affect measure, study design, participants’ age, year of publication, and whether the study insured that only pre-consumption affect was measured, were examined in the analysis. In light of literature highlighting different publications practices of outlets (81), journal was another variable included as a moderator. Country and gender were also included in moderator analysis (following 18,46).

Method

Operational definitions

This review focused on daily alcohol consumption likelihood, which is defined as the odds of one’s alcohol consumption on a given day. Mood and emotions are distinct but interrelated constructs in that the former tend to be more stable and “flat”, while the latter are construed as more vivid and quick (82). However, studies sometimes use these terms interchangeably. Therefore, to account for differences in the terminology, the terms ‘affect’, and ‘affective state’ are used in this review as umbrella terms for the experience of mood, emotion, or feelings.

Transparency and openness

We report how eligible studies were selected, the databases that were searched, what data were extracted, and the process of data analyses. Files used for data analysis, as well as the analysis code, can be found on the Open Science Framework: https://osf.io/cp86f/. The study design and its analysis were not pre-registered.

Eligibility criteria

The literature search and eligibility criteria were developed prior to conducting the systematic review. The relevance of each study was assessed according to the following inclusion criteria: peer-reviewed
papers; focus on the general human population (non-clinical sample); looking at affective states on the day of consumption; looking at consumption likelihood or odds of drinking (rather than alcohol consumption volume, rate of consumptions, blood alcohol concentration, or alcohol craving); conducting odds ratio analysis (or having statistics from which odds ratios along with the confidence intervals could be extracted); papers in English or Russian. The exclusion criteria were as follows: reviews, books, posters, and editorials; literature examining clinical samples (individuals with alcohol use disorders or any other clinical disorder). As there were no previous meta-analyses on the topic, the year range was not specified.

Studies examining positive and/or negative affect were included, as were those that measured affect on a continuum (i.e., where positive and negative affect are at polar ends of the same assessment spectrum) or as separate entities. Furthermore, studies that examined mean levels of affect (i.e., average negative or positive affect) as well as affect facets (i.e., specific emotions, e.g., stress, anger, or happiness) were included.

**Literature review**

The literature search was primarily conducted by the lead author. A comprehensive search was conducted of the following databases: PsychINFO, PsycARTICLES, Science Direct, Wiley Online Library, PubMed, SCOPUS, JSTOR using Preferred Reporting Items for Systematic Review and Meta-Analyses (83) and American Psychological Association’s Meta-Analysis Reporting Standards (84,85) methodologies. The search terms were developed by primary author based on keywords used by relevant articles and were discussed and refined with the rest of the team prior to the start of data collection. The following commands were used for searching: (“alcohol consumption likelihood” OR “drinking likelihood” OR “odds of drinking” OR “odds of alcohol consumption”) AND (“mood” OR “emotion” OR “affective states”) NOT “disorders”. After the initial literature search, to avoid missing data, the second author conducted a comparative title search using the same search terms and eligibility criteria to ensure the incorporation of any studies that may have been overlooked in the original review. Bibliographies from relevant reviews and book chapters, as well as articles that fit the inclusion criteria, were manually searched for additional citations. Full-text papers of any titles and abstracts that were considered relevant were obtained where possible.

**Quality assessment and data extraction**

Study quality was assessed using standard criteria (86), with papers screened by two independent reviewers (Cohen’s Kappa = .71). None of the studies were judged to be of poor quality (and hence none were excluded), while three were deemed to be of moderate quality, and eight were classified as representing good quality.

Following the quality assessment, relevant data were extracted from each study (see Table 1 for a full summary). The data were extracted by one of the authors using the Excel template (extracted data summarized in the table are available on the Open Science Framework) and subsequently checked by the first author. All effect sizes pooled from the studies were on the daily level of analysis. To examine how affect may impact subsequent drinking likelihood, effects of affect at time-1 were extracted from the papers. Some data sets provided multiple odds ratios for the constructs of interest (e.g., odds of drinking following sadness and following anger, both of which relate to the negative affect). Common ways of dealing with this problem include choosing a single outcome (87) or aggregating all measures by computing an average effect size (88); however, such approaches may lose information, rule out critical analyses of within-study factors (e.g., informant differences), artificially reduce the variance between effect sizes, and risk inaccurate estimation of study effect sizes (89–91). Therefore, we used the robust variance estimation (92) method to control for dependencies between effect sizes.

**Meta-analysis - analytical strategy**

Pre-calculated odds ratios, along with their corresponding 95% confidence intervals, were used as effect sizes for meta-analyses. Meta-analyses on positive and negative affect were performed separately.

Analysis was conducted in R studio (93) using the robumeta package (94). First, odds ratios and their corresponding confidence intervals were log-transformed. Then, the standard error was calculated by extracting log-transformed lower confidence interval from the upper confidence interval and dividing it by 3.92. After that, the variance was calculated by squaring standard error. The intercept-only model was fitted using the robu function. Heterogeneity of the studies was assessed using I² statistics. Because correlations between the effect sizes reported within each study were not known, we assumed a Spearman’s rho (ρ) of .80 (92). We also performed a series of sensitivity analyses by testing different values of ρ in intervals of .10. This did not affect inferences about effect sizes; therefore, these results are not reported in the
### Table 1. Summary of the studies.

| Number | Authors | Year | Country | Length of the study | Number of assessments per day | Sample size(s) | Sample type | Sample gender(s) | Sample age(s) | Positive or negative | Affect measure |
|--------|---------|------|---------|--------------------|------------------------------|----------------|-------------|-----------------|----------------|---------------------|----------------|
| 1      | Duif, Thewissen, Wouters, Lechner, & Jacobs | 2019 | Netherlands | 7                | 11                           | 162 adults     | Adults      | 109 women       | M = 36.07, SD = 9.27, range = 20 - 50 | Both | PANAS Dvorak & Simons |
| 2      | Gottfredson & Hussong | 2014 | USA | 21 | 9 | 100 young adults | 61 women | M = 20.09, SD = 1.67, range = 18 - 25 | Both | PANAS (Watson et al., 1988) and mood circumplex (Diener et al., 1983) |
| 3      | Shadur, Hussong, & Haroon | 2013 | USA | 28 | 1 | 86 young adults | 46 women | M = 18, SD not reported, range = 18-20 | Both | PANAS-X (Watson & Clark, 1994) |
| 4      | Feagans Gould, Hussong, & Hersh | 2013 | USA | 21 | 4 | 73 adolescents | 48 women | M = 13.92, SD = .47, range not reported | Negative | Sadness and worry on a scale from 1 to 5 |
| 5      | Jones, Allen, Lanza, Graham-Engeland | 2012 | USA | 21 | 4 | 75 adolescents | 49 women | M = 13.6, SD & range not reported | Negative | MAACL-R (Lubin et al., 1998) |
| 6      | Luk, Fairlie, & Lee | 2021 | USA | 7 | 9 | 93 adults | 58 women | M & SD not reported, range = 25 - 65 | Both | Circumplex model of affect (Posner et al., 2005; Russell, 1980) |
| 7      | Patrick, Yeomans-Maldonado, & Griffin | 2018 | USA | 14 | 3 | 347 young adults | 186 women | M = 19.7, SD = 1.3, range not reported | Negative | PANAS (Watson et al., 1988) |
| 8      | Stamates, Linden-Carmichael, Preonas, & Lau-Barraco | 2016 | USA | 14 | 1 | 72 young adults | 20 women | M = 18.7, SD and range not reported | Both | PANAS (Watson et al., 1988) |
| 9      | Stamates, Linden-Carmichael, Preonas, & Lau-Barraco | 2019 | USA | 14 | 1 | 24 young adults | 14 women | M = 23.83, SD = 1.83, range not reported | Both | PANAS (Watson et al., 1988) |
| 10     | O'Hara, Armeli, & Tennen | 2019 | Australia | 21 | 3 | 83 young adults | 63 women | M = 21.42, SD = 3.09, range = 18 - 30 | Both | Happy, relaxed, irritated, stressed, on a 6-point scale |
| 11     | Dvorak, Pearson, Sargent, Stevenson, & Mfon | 2014 | USA | 30 | 1 | 1636 young adults | 867 women | M = 19.2, SD = 1.4, range not reported | Both | Several adjectives from PANAS-X (Watson & Clark, 1994) |
| 12     | Dvorak, Pearson, Sargent, Stevenson, & Mfon | 2016 | USA | 21 | 3 | 74 university students | 43 women | M = 21.30, SD & range not reported | Both | PANAS-X (Watson & Clark, 1994) and mood circumplex (Larsen & Diener, 1992) |
paper. Using transf.exp.int function, the results were transformed back from the logs to odds ratios. Odds ratios greater than one indicated a positive association, odds ratios below one indicated a negative association.

Multiple moderators were examined, including journal, country, sample age (adolescents, young adults, adults), affect measure, the length of the study, the number of assessments per day, and whether the study controlled for the assessment of affect prior to drinking occasion. When analyzing affect measure as a moderator, due to heterogeneity between the measures (e.g., using different items from PANAS, using a combination of PANAS with other measures), studies were classified into following categories: PANAS, Circumplex Model of Affect, and MAACL.

As the robustmeta package does not support the function of building forest plots for odds ratios, cluster-robust models were fitted using the metafor (95) and clubSandwich (96) packages. Using transf.exp.int function, the results were transformed back from the logs. The forest plots were built using back-transformed odds ratios. Power of meta-analyses was calculated using an existing formula (97).

In case of significant results, publication bias was examined via Egger’s test, which was performed by imputing standard error as a moderator to the main model. Additionally, the three-parameter selection model based on step functions was used to estimate the potential publication bias using the selmodel command. As the selmodel only accepts models created by metaphor’s rma function, the model was first fitted using this function. The cut-points of .025 was chosen, and, as a sensitivity analysis, a cutoff point of .05 was tested as well. To perform a p-curve analysis and assess potential p-hacking (98), single-level models were used (computed using metagen function from meta package; (99)). Similarly, to find outliers, find.outliers function was used on a single-level model. We also computed a fail-safe N, the number of additional “zero-effect” studies needed to increase the p-value for the meta-analysis to above .05 (100). Sensitivity analysis was conducted to account for studies that were not included in meta-analyses. To account for family-wise error, Bonferroni corrections were applied to significant results.

**Results**

**Quantity of research available**

Electronic and hand searches identified 15,374 articles, which, once duplicates were removed, yielded 2474 unique citations to be screened for inclusion (Figure 1). Their titles and abstracts were assessed for their relevance to the review, resulting in 21 potential articles being retained. The full texts of all these studies were obtained. After applying exclusion criteria for the remaining full-text papers, 14 articles were excluded. After that, full texts of eligible articles were screened to obtain additional citations. This resulted in screening five additional articles, all of which were retrieved and considered eligible. Overall, 13 studies were eligible for systematic review, 11 were eligible for the negative affect meta-analysis, and nine were eligible for the positive affect meta-analysis. The PRISMA flow diagram summarizes the included studies for both negative and positive affect (see Figure 1). The main reason for the exclusion of studies was that they did not look at odds of alcohol consumption as an outcome.

Included studies were published between 2012 and 2021 and were conducted in various countries: Australia (n = 1), Canada (n = 1), Netherlands (n = 1), and the USA (n = 10). Two studies sampled adolescents, seven studies sampled young adults, and two studies used adult samples. All studies utilized intensive longitudinal design methods. The number of assessments per day ranged from between one and 11, and the length of the studies varied from 7 to 30 days. Study characteristics are provided in Table 1, and a summary of effect sizes is provided in Table 2.

**Association between daily negative affect and drinking likelihood**

A total of 11 studies and 2751 participants are represented in the analysis. Odds ratio ranged from .12 to 1.98. Analysis revealed high levels of heterogeneity, I^2 = 82.34. The log-transformed effect size coefficient was β = -.10, CI [-.32, .12], SE = .10, t = -.09, p = .309. When the coefficient was transformed back into odds ratio, the result was: OR = .90, CI [.73, 1.12], indicating that there is no significant association between same-day negative affect and odds of alcohol consumption. The power of this meta-analysis was high (.88, assuming high heterogeneity). Figure 2 shows the sample effect size in the forest plot.

**Moderation analysis**

None of the moderators had a significant effect on study results. See supplementary materials for non-significant moderation findings.

**Outliers**

Four outliers were identified in a random-effects model (28,57,101,102). When analyzed without outliers (using single-level model), heterogeneity decreased (I^2 = 58.2), but the results of meta-analysis did not significantly change, OR = .96, CI [.89, 1.06], p = .485, as negative affect was still not significantly associated with consumption.
Sensitivity analysis

Sensitivity analysis was conducted, which included two additional samples, for which effect sizes were imputed as 0.0. This analysis yielded a weighted mean effect size of estimate = −.05, 95% CI [−.33, .24], p = .734. This indicates that the missing data likely had little effect on the estimate of the effect size and significance. The results of the sensitivity analysis did not reveal significant differences with different values of Rho. Consequently, these results are not reported in the paper.

Association between daily positive affect and drinking likelihood

A total of 8 studies and 2244 participants are represented in the analysis. Odds ratio ranged from .19 to 1.52. Analysis revealed moderate levels of heterogeneity, $I^2 = 54.97$. The log-transformed effect size coefficient was $\beta = .16$, CI [.09, .24], SE = .03, $t = 5.63$, $p = .003$. When the coefficient was transformed back into odds ratio, the result was: OR = 1.17, CI [1.09, 1.27], indicating that there is a significant association between same-day positive affect and odds of alcohol consumption. When Bonferroni correction was applied (updated $p$ value of .005), the results remained significant. The power of this meta-analysis was high (.99, assuming moderate heterogeneity). Figure 3 shows the sample effect size in the forest plot.

Moderation analysis

Country was the only significant moderator, but as degrees of freedom were lower than four, the estimate could not be trusted. See supplementary materials for the summary of results for non-significant moderators.

Publication bias

Egger’s test did not find significant results for publication bias (see summary of the model in the supplementary materials). $P$-curve analysis indicated that
Table 2. Effect size summary.

| Authors                                      | OR  | Lower CI | Upper CI | OR  | Lower CI | Upper CI |
|----------------------------------------------|-----|----------|----------|-----|----------|----------|
| Duff, Thewissen, Wouters, Lechner, & Jacobs (2019) | 1.98 | 1.11     | 3.53     | 1.33 | 0.90     | 1.98     |
| Duff, Thewissen, Wouters, Lechner, & Jacobs (2019) | 1.11 | 0.92     | 1.35     | 1.21 | 1.05     | 1.40     |
| Duff, Thewissen, Wouters, Lechner, & Jacobs (2019) | 0.89 | 0.56     | 1.42     | 1.00 | 0.74     | 1.35     |
| Duff, Thewissen, Wouters, Lechner, & Jacobs (2019) | 0.99 | 0.75     | 1.30     | 0.96 | 0.80     | 1.14     |
| Dvorak & Simons (2014)                        | 1.01 | 0.88     | 1.15     | 1.20 | 0.95     | 1.22     |
| Gottfredson & Hussong (2013)                  | 1.08 | 0.91     | 1.18     |      |          |          |
| Shadur, Hussong, & Haroon (2015)              | 0.93 | 0.69     | 1.26     |      |          |          |
| Shadur, Hussong, & Haroon (2015)              | 0.93 | 0.70     | 1.25     |      |          |          |
| Shadur, Hussong, & Haroon (2015)              | 1.27 | 0.94     | 1.72     |      |          |          |
| Shadur, Hussong, & Haroon (2015)              | 1.35 | 1.01     | 1.84     |      |          |          |
| Feagans, Gould, Hussong, & Hersh (2012)       | 0.12 | 0.05     | 2.91     |      |          |          |
| Jones, Allen, Lanza, Graham-Engel (2021)      | 0.83 | 0.47     | 1.45     | 0.92 | 0.64     | 1.32     |
| Jones, Allen, Lanza, Graham-Engel (2021)      | 1.31 | 0.79     | 2.18     | 1.52 | 1.12     | 2.09     |
| Luk, Fairlie, & Lee (2018)                    | 0.65 | 0.59     | 1.16     |      |          |          |
| Luk, Fairlie, & Lee (2018)                    | 0.32 | 0.37     | 1.00     |      |          |          |
| Patrick, Yeomans-Maldonado, & Griffin (2016)  | 0.65 | 0.53     | 0.81     | 1.34 | 1.20     | 1.50     |
| Patrick, Yeomans-Maldonado, & Griffin (2016)  | 0.32 | 0.18     | 0.56     | 1.19 | 0.86     | 1.65     |
| Stamates, Linden-Carmichael, Preonas, & Lau-Barraco (2019) | 0.90 | 0.77     | 1.05     | 1.08 | 1.00     | 1.16     |
| O'Donnell et al. (2019)                       | 0.90 | 0.75     | 1.22     | 0.19 | 0.10     | 2.41     |
| O'Donnell et al. (2019)                       | 1.03 | 0.76     | 1.35     | 0.97 | 0.76     | 1.28     |
| O'Hara, Armeli, & Tennen (2014)               | 1.05 | 0.98     | 1.06     | 1.26 | 1.00     | 1.22     |
| O'Hara, Armeli, & Tennen (2014)               | 1.08 | 0.97     | 1.10     |      |          |          |
| O'Hara, Armeli, & Tennen (2014)               | 1.26 | 1.00     | 1.22     |      |          |          |
| Dvorak, Pearson, Sargent, Stevenson, & Mfon (2016) | -   | -        | -        | 1.19 | 1.00     | 1.16     |

The evidential value is present, and that evidential value is not absent or inadequate (see Figure 4). This means that P-curve estimates that there is a “true” effect size underlying finding, and that the results are unlikely to be the product of publication bias and p-hacking alone.

In addition, the fail-safe N showed that 192 studies with OR = .00 would need to be incorporated into the meta-analysis to yield a non-significant effect. This markedly exceeded the benchmark of 115 (5n + 10) (100), suggesting that our findings are robust to the threat that excluded studies might have yielded a non-significant effect.

The .025 selection model’s estimate of the true average effect size was .105, CI (95% CI [−.02 to .23]). When converted back to OR, OR = 1.110, 95% CI [.98, 1.26]. As the test of selection parameters were non-significant ($\chi^2 = 1.40, p = .236$), this indicates that the meta-analysis was not substantially biased by a lower selection probability of non-significant results. The analysis on .05 level also indicated that there was no substantial publication bias in the studies (logOR = .15, CI [.05, .24], when converted back OR = 1.16 [1.05, 1.28]).

Sensitivity analysis

Sensitivity analysis was conducted, which included four additional samples, for which effect sizes were imputed as 0.0. This analysis yielded a weighted mean effect size of estimate = .15, 95% CI [.08, .22], $p < .001$; OR = 1.16, CI [1.08, 1.25]. This indicates that the missing data likely had little effect on the estimate of the effect size and significance. The results of the sensitivity analysis did not reveal significant differences with different values of Rho; therefore, these results are not reported in the paper.

Discussion

Building on previous theoretical and empirical work (13,46–48,103), the aim of this study was to clarify and consolidate the literature on the impact of intra-day affective states on alcohol consumption likelihood in non-clinical samples. Since there is evidence that positive and negative affect are distinct processes, two separate meta-analytic reviews were conducted on the impact of negative (review one) and positive (review two) affective states on daily alcohol consumption likelihood. Contrary to the self-medication hypothesis (13), results indicate that negative affect was not associated
with alcohol consumption likelihood, while a significant but weak association between increased positive affect and alcohol consumption likelihood was found. Present results were consistent across different age groups (adolescents, adults, and young adults), study design (number of days/assessments per day) and independent of affect conceptualization (i.e., scales used).

A reason for the association between positive (and not negative) affect and alcohol consumption likelihood may lie in the fact that alcohol is often used for celebratory occasions (3), while, depending on personal characteristics, alternative strategies of coping with distress may be more prevalent in non-clinical samples (104). In conjunction with the consolidated evidence base regarding daily affective states and alcohol consumption volume (46), the current findings shed further light on what appears to be a complex relationship between affect and alcohol use. Specifically, recent meta-analyses suggest that experimentally induced negative affect is related to increased alcohol consumption volume and craving (47), while elevated affect intensity (in both lab and field contexts) seems to be associated with increased alcohol consumption volume (46). From the results of the current study, it appears that positive but not negative affect is associated with drinking likelihood on a given day. Seen in relation to the findings of previous meta-analyses in this field of research, this may therefore suggest that when drinking has already started, or if a drinking occasion is planned on a day...
characterized by increased negative affect, then it may be that affect intensity regulation, as opposed to valence, contributes to increased drinking volume (46). That could explain the discrepancy between the findings of the current analyses and previous ones (46,47). From this perspective, people may not necessarily initiate alcohol consumption to deal with negative affect but such moods may contribute to elevated consumption volume if alcohol is readily available, or consumption has already commenced. On the other hand, people appear more likely to consume alcohol and drink more when they experience elevated positive mood (i.e., for enhancement drinking motives). As drinking alcohol is a normative way of celebrating in many Western societies (105), individuals who are in a good mood, and want to celebrate, may be more likely to start drinking and to drink increased amounts. Further studies on whether the same holds true for clinical populations appear necessary, as a previous meta-analysis on emotion regulation and substance use points to larger effect sizes in clinical samples (18), although a recent pre-print (106) indicates that the same relationship between daily affect and alcohol consumption likelihood, observed in the current study, also holds in such samples.

Combined with results of previous studies, our findings could help to inform prevention efforts. As such, interventions for people who wish to reduce their alcohol consumption may focus on days characterized by increased positive affect (and/or planned celebratory occasions) and encourage individuals to engage in

Figure 3. Forest plot of studies examining the relationship between daily negative affect and odds of alcohol consumption. Increased positive affect is associated with higher odds of drinking.
ways of elevating further their mood that do not involve alcohol consumption. On the other hand, for days characterized by increased levels of negative affect, such efforts may only be necessary if individuals plan drinking occasions, or when alcohol is readily available to them.

Considering study limitations, it is important to note that, while the power of our overall meta-analyses was high, it may still have been insufficiently powered to examine the effects of various moderators due to the number of available studies and their heterogeneity. It must also be noted that, owing to the inconsistent way data on socioeconomic status were (not) collected within the extant literature, it was not possible to conduct moderation analyses on this factor. Additionally, while the search terms were carefully selected and based on keywords utilized in the extant literature, no consultation with field experts or journal indexers was conducted, raising a potential concern of selection bias. Furthermore, this study was only concerned with drinking likelihood. That is, we looked at the odds of alcohol consumption on a given day. While there are meta-analytical studies that are concerned with drinking volume (46) and alcohol cravings (47), in order to fully understand the nature of the relationship between affect and alcohol consumption, future research is advised to examine other variables of interest such as blood alcohol concentration or rate of consumption. Furthermore, this review only focused on intra-day consumption. While this allowed us to examine the association between state affect, further examination of trait affect (i.e., tendency to experience particular affective states) could help answer the question of how longer-term affective states may be associated with alcohol consumption (107). Similarly, not all studies made it clear whether they controlled for the temporal order of affect and alcohol consumption, and, although moderator analysis revealed no difference between the studies with regard to this, the results of some studies may have been biased by not controlling for the temporal proceedings of affect. More generally, there is also a need for studies to be adequately powered and to conduct longitudinal investigations given the dominance of cross-sectional work in this area.

We conclude that meta-analytic findings converge to suggest that increased positive affect is weakly associated with greater consumption likelihood in non-clinical populations, supporting the positive reinforcement model of alcohol use. The self-medication hypothesis was not supported as far as drinking likelihood as the dependent variable of interest is concerned. Combined with the contributions of previous consolidatory studies, the results
of the present meta-analyses suggest that affective states and their regulation have a notable and nuanced influence on alcohol consumption behaviors.

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