Harris-Fry, HA; Paudel, P; Harrisson, T; Shrestha, N; Jha, S; Beard, BJ; Copas, A; Shrestha, BP; Manandhar, DS; Costello, AML; Cortina-Borja, M; Saville, NM (2018) Participatory Women's Groups with Cash Transfers Can Increase Dietary Diversity and Micronutrient Adequacy during Pregnancy, whereas Women’s Groups with Food Transfers Can Increase Equity in Intrahousehold Energy Allocation. The Journal of nutrition. ISSN 0022-3166 DOI: https://doi.org/10.1093/jn/nxy109

Downloaded from: http://researchonline.lshtm.ac.uk/4648717/

DOI: 10.1093/jn/nxy109

Usage Guidelines

Please refer to usage guidelines at http://researchonline.lshtm.ac.uk/policies.html or alternatively contact researchonline@lshtm.ac.uk.

Available under license: http://creativecommons.org/licenses/by/2.5/
Participatory Women’s Groups with Cash Transfers Can Increase Dietary Diversity and Micronutrient Adequacy during Pregnancy, whereas Women’s Groups with Food Transfers Can Increase Equity in Intrahousehold Energy Allocation

Helen A Harris-Fry,1 Puskar Paudel,2 Tom Harrisson,3 Niva Shrestha,1 Sonali Jha,2 B James Beard,1 Andrew Copas,4 Bhim P Shrestha,2 Dharma S Manandhar,2 Anthony M de L Costello,3 Mario Cortina-Borja,4 and Naomi M Saville1

1London School of Hygiene & Tropical Medicine, London, United Kingdom; 2Mother and Infant Research Activities (MIRA), Kathmandu, Nepal; 3Institute for Global Health; and 4Great Ormond Street Institute of Child Health, University College London, London, United Kingdom

Abstract

Background: There is scarce evidence on the impacts of food transfers, cash transfers, or women’s groups on food sharing, dietary intakes, or nutrition during pregnancy, when nutritional needs are elevated.

Objective: This study measured the effects of 3 pregnancy-focused nutrition interventions on intrahousehold food allocation, dietary adequacy, and maternal nutritional status in Nepal.

Methods: Interventions tested in a cluster-randomized controlled trial (ISRCTN 75964374) were “Participatory Learning and Action” (PLA) monthly women’s groups, PLA with transfers of 10 kg fortified flour (“Super Cereal”), and PLA plus transfers of 750 Nepalese rupees (∼US$7.5) to pregnant women. Control clusters received usual government services. Primary outcomes were Relative Dietary Energy Adequacy Ratios (RDEARs) between pregnant women and male household heads and pregnant women and their mothers-in-law. Diets were measured by repeated 24-h dietary recalls.

Results: Relative to control, RDEARs between pregnant women and their mothers-in-law were 12% higher in the PLA plus food arm (log-RDEAR coefficient = 0.12; 95% CI: 0.02, 0.21; P = 0.014), but 10% lower in the PLA-only arm between pregnant women and male household heads (−0.11; 95% CI: −0.19, −0.02; P = 0.020). In all interventions, pregnant women’s energy intakes did not improve, but odds of pregnant women consuming iron-folate supplements were 2.5–4.6 times higher, odds of pregnant women consuming more animal-source foods than the household head were 1.7–2.4 times higher, and midupper arm circumference was higher relative to control. Dietary diversity was 0.4 food groups higher in the PLA plus cash arm than in the control arm.

Conclusions: All interventions improved maternal diets and nutritional status in pregnancy. PLA women’s groups with food transfers increased equity in energy allocation, whereas PLA with cash improved dietary diversity. PLA alone improved diets, but effects were mixed. Scale-up of these interventions in marginalized populations is a policy option, but researchers should find ways to increase adherence to interventions. This trial was registered at www.controlled-trials.com as ISRCTN 75964374. J Nutr 2018;148:1472–1483.

Keywords: pregnancy, dietary adequacy, nutrition, food allocation, supplements, cash transfers, women’s groups, community interventions, Nepal, maternal health

© 2018 American Society for Nutrition. This is an Open Access article distributed under the terms of the Creative Commons Attribution License (http://creativecommons.org/licenses/by/4.0/), which permits unrestricted reuse, distribution, and reproduction in any medium, provided the original work is properly cited.
Interventions to improve diets and food allocation

Interventions (3). Interventions such as food supplements, from this region on the effectiveness of common nutrition poorest in the world (1, 2), but there is a striking lack of evidence this region against women in the allocation of food is more pervasive in this region than elsewhere (11, 12), particularly during pregnancy (13), and this may prevent women from consuming transfers delivered at the household level (14). Therefore, trials from other locations may not be valid for South Asia, and interventions may need to incorporate additional components to change intrahousehold allocation practices (14). Common approaches to change these food-related behaviors include nutrition education, counseling, and women's group interventions (15–17). One potential intervention is women's groups using a “Participatory Learning and Action” (PLA) approach (18), in which group members discuss and prioritize health issues, identify local barriers to good health, and implement and informally evaluate strategies to address these barriers (16, 19). A meta-analysis found that PLA women's groups reduced neonatal mortality by 20%, and by 33% when >30% of the group members were pregnant (20), but impacts on nutritional outcomes are weaker (21, 22). To better improve nutritional outcomes, women's group members may require additional resources, such as food or cash transfers, so they can act upon new knowledge and skills.

The Low Birth Weight South Asia Trial (LBWSAT) tested the effects of monthly PLA women’s groups, with or without cash [750 Nepalese rupees (NPR); US$7.5/mo] or food (10 kg fortified wheat-soya “Super Cereal”/mo) transfers during pregnancy in Nepal, on birth weight (≤72 h after birth) and weight-for-age z scores (children aged 0–16 mo) (23). Birth weight improved by 78 g (95% CI: 13.9, 142 g) in the PLA plus food arm but did not significantly improve in the PLA-only or PLA plus cash arms (24), nor were there any effects on weight-for-age in any arms. This substudy, nested within LBWSAT, aimed to understand how these interventions influenced dietary patterns and whether they selectively targeted pregnant women, by measuring the effects on maternal dietary adequacy, intrahousehold food allocation, and nutritional status during pregnancy.

Methods

Study setting and population

The trial (ISRCTN 75964374) was implemented in the Dhanusha and Mahottari districts (in province 2) in Nepal by University College London and a Nepalese research organization, Mother and Infant Research Activities (MIRA), in collaboration with the World Food Programme and Save the Children UK. These districts were purposively sampled because the burden of undernutrition is high: 28% of infants are born with a low birth weight (25) and 29% of women are underweight (26).

Dhanusha and Mahottari districts have a combined population of ∼1.4 million (27). Flooding is frequent during the monsoon, and poor-quality roads, high temperatures, and humidity make travel to study clusters difficult. The study period coincided with the aftermath of the 2015 earthquakes and severe political unrest associated with Nepal’s new constitution (28). This caused strikes, violent protests, road blockages, closure of markets and banks, and personal insecurity for the field team, making data collection difficult.

Random allocation

The interventions were community based, so were allocated by cluster. Each cluster was defined as 1 Village Development Committee, a geopolitical unit that contains 9 wards. Clusters were excluded if they had been allocated to a women’s group intervention in an earlier trial. To reduce the heterogeneity of the sample, clusters were also excluded if they were small (population <4000), large (population >9200), near to the East-West Highway, hilly, forested, non–Maithili-speaking, or large towns or municipalities. In a public “bingo” event involving local stakeholders, the remaining 80 clusters were allocated to 1 of 4 trial arms, giving 20 clusters per arm. We used stratified block randomization. Strata were based on cluster size and accessibility: small inaccessible, small accessible, large inaccessible, and large accessible. Due to the nature of the interventions, masking of the cluster allocations was not possible. Statisticians were not masked to the allocation analyses followed a prespecified plan.

Participants

All pregnant women living in selected clusters were eligible to participate in the trial, but this study evaluated the effects of the interventions on diets and intrahousehold food allocation in joint, male-headed households. This was to reduce the heterogeneity of the sample and because these households were hypothesized to be the least equitable and least likely to change their behaviors (14). Criteria for enrolled women and their households to be interviewed for this study were as follows: women in their third trimester who were living with their mother-in-law in male-headed households. Sampled individuals from these households were pregnant women, their mothers-in-law, and male household heads. Male household heads were hypothesized to be the most favored in terms of food allocation, and mothers-in-law typically controlled food allocation (13). We initially planned to collect data over 1 y but were delayed due to severe political conflict, major earthquakes, and lack of funds.

Interventions

The full protocol is reported in Saville et al. (23). The 3 interventions were a behavior-change strategy of monthly women’s groups practicing...
PLA, PLA women’s groups with a monthly unconditional food transfer, and PLA women’s groups with a monthly unconditional cash transfer.

The groups followed a PLA cycle of 4 phases: identify problems, plan strategies, act together, and evaluate impact (as shown in Figure 1).

In the first phase, groups used participatory methods such as picture cards, games, and stories to discuss nutrition problems and local barriers to achieving good health during pregnancy. In the second phase, groups prioritized and voted on the issues they wanted to focus on, designed strategies to address these problems, and engaged the wider local community for support and feedback. In the third phase, the groups implemented these strategies while continually discussing new topics related to pregnancy and infant health. Finally, in the fourth phase, the groups reviewed what went well and discussed what to do next after the implementing organization (MIRA) withdrew from the community.

The women’s groups were held at the ward level and, with 9 wards per Village Development Committee, 539 women’s groups were facilitated in all 60 intervention clusters (one small ward was merged with a neighboring ward). Government-incentivized female community health volunteers (FCHVs) facilitated the PLA women’s groups, with support from MIRA-recruited “nutrition mobilizers.” FCHVs and nutrition mobilizers received training on group facilitation and maternal and infant nutrition.

The PLA groups served as a mechanism to deliver the food and cash transfers. The food consisted of 10 kg of fortified wheat and soya blend with sugar (“Super Cereal”)—a product developed and tested by the World Food Programme—and the cash was NPR 750 (23). The nutritional composition and safety of Super Cereal is reported in the trial protocol (23). Women were advised to eat 150 g Super Cereal/d, which would provide 570 kcal/d and meet most micronutrient requirements for pregnant women. The monthly 10-kg transfer allowed for women to share another 179 g/d with other household members. The cash transfer was calculated to be the equivalent cost of the Super Cereal transfer. The cost of an adequate diet in 2012 was estimated at NPR 1767/mo per person, so the cash transfer would provide 43% of this cost. In the PLA plus food arm, PLA groups discussed how much Super Cereal should be consumed (150 g/d) and recipes for using the Super Cereal. In the PLA plus cash arm, groups discussed how the cash should be spent on nutritious foods such as milk and fruit. In both transfer arms, groups discussed why the transfers should be selectively consumed by pregnant women.

As a benefit to all participants (including controls), training on maternal, newborn, and child health was provided to 189 health workers in both districts. When the final measurements were taken after delivery, women in the control and PLA-only arms received NPR 1000 to thank them for participating.

Outcomes
The outcomes are summarized in Table 1. We report intervention effects on food and nutrient intakes of pregnant women and intrahousehold allocation. To describe both intakes and allocation, we selected nutrients (energy, protein, and dietary iron) and food groups (“flesh foods”, i.e., meat, fish, or shellfish; dairy foods; and green leafy vegetables) that were promoted in the PLA groups. Other foods were also promoted, but we chose flesh foods because they are culturally high status, dairy foods because they are high-status foods that are cheaper and more regularly consumed than flesh foods, and green leafy vegetables because they are micronutrient-rich yet considered “inferior.” We report on intakes of dietary iron (excluding iron-folate supplements) and consumption of supplements, and on intrahousehold allocation of dietary and total iron (dietary iron plus iron from supplements). Midupper arm circumference (MUAC) is reported as an indicator of nutritional status.

We also report summary measures that were added to the initial analysis plan to describe overall nutritional and dietary adequacy: mean probability of adequacy (MPA), the average adequacy of 11 nutrients, and the Minimum Dietary Diversity Score for Women (29), the number of food groups consumed out of 10 food groups.

The primary outcomes for this study were mean Relative Dietary Energy Adequacy Ratio (RDEAR) between pregnant women and the
male household heads and mean RDEARs between pregnant women and their mothers-in-law, as defined in the following equation:

\[
\text{RDEAR} = \frac{\frac{kcal \text{ intake}_a}{kcal \text{ Estimated Average Requirements}_a}}{\frac{kcal \text{ intake}_b}{kcal \text{ Estimated Average Requirements}_b}};
\]

where \(a\) denotes the pregnant woman and \(b\) denotes another household member (the male household head or mother-in-law). Estimated Average Requirements (EARs) were individually calculated, on the basis of their age, sex, body weight, pregnancy/lactating status, and self-reported physical activity levels (30). Total energy expenditure was calculated as basal metabolic rate \(\times\) physical activity levels. Basal metabolic rates were based on Indian-specific values (31). Self-reported activity levels were assessed by asking each respondent whether they spent most of the previous day 1) sitting down or standing still, such as doing office or shop work (“sedentary”); 2) moving around, such as walking, carrying light loads, or doing domestic work (“moderate”); or 3) doing strenuous work like carrying heavy loads, working in the field, or pulling a rickshaw (“strenuous”). Based on Indian Council of Medical Research’s guidelines, Physical activity levels were taken as 1.5 for sedentary activity, 1.8 for moderate work, and 2.3 for strenuous, heavy activity levels (31). The additional cost of pregnancy was taken to be 390 kcal/d (31).

We calculated the ratios of the MPA between pairs of household members. MPA was calculated by using the Probability approach (32), as follows. We first transformed nutrient intake distributions to normality using Box-Cox transformations (33) and obtained the best linear unbiased predictors derived from separate linear mixed-effects regression models for each household member, treating clusters and individuals (repeat measurements) as random effects and strata and study arm as fixed effects. We then calculated the probability of adequacy (PA) for each individual and each nutrient by comparing the individual’s back-transformed usual intake to the population distribution of requirements, which are normal distributions with known means (i.e., EARs) and SDs (34–37). For mothers-in-law and male household heads, iron PAs were calculated by using a table of probabilities for different intervals of usual intakes (36). We assumed low bioavailability of iron (5%), apart from pregnant women who have higher iron absorption (23%) (36), and low bioavailability of zinc (25% for women, 18% for men) (37).

We also report, but did not formally test, the percentage of women with a probability of iron adequacy of \(>0.7\) based on total iron including supplements.

### Field procedures

Enrollment of pregnant women into the main trial ran from December 2013 to February 2015, and the interventions ran from February 2014 to October 2015. The last enrolled woman gave birth in October 2015. Between 10 June and 26 September 2015, all eligible households were visited for a series of three 24-h dietary recall interviews. Dietary intakes...
of pregnant women, male household heads, and mothers-in-law were assessed up to 3 times each on nonconsecutive days. The 5-stage “multi-pass” dietary assessment method has been described in detail elsewhere (38–40). Portion sizes were estimated by using an atlas of life-sized photographs that was developed locally and validated before use (41).

Individual nutritional intakes were estimated with the use of a local food-composition table compiled from multiple published sources (42–45). We also added the nutritional content of cooked, mixed dishes from a local food-composition table compiled from multiple published sources (38–40). Portion sizes were estimated by using an atlas of life-sized photographs that was developed locally and validated before use (41).

MUAC and body weight were measured with the use of Seca 212 circumference tapes and Tanita solar weighing scales accurate to 100 g. Socioeconomic and demographic data were collected by the main trial surveillance (23, 46). Data on exposure to the interventions were assessed up to 3 times each on nonconsecutive days. The 5-stage “multi-pass” dietary assessment method has been described in detail elsewhere (38–40). Portion sizes were estimated by using an atlas of life-sized photographs that was developed locally and validated before use (41).

Before data collection, interviewers received standardization exercises for MUAC measurement and 1 wk of training on dietary assessment. To ensure data quality, supervisors and project coordinators conducted spot checks, revisited interviewed households, and checked plausibility of values, frequency of outliers, digit preference, and Global Positioning System points every week.

Ethics
Research ethics approval was obtained from the Nepal Health Research Council (108/2012) and the University College London Ethical Review Committee (4198/001). Women gave consent by signature or thumbprint. A data-monitoring committee was formed, and harms were tracked in the form of stillbirths and deaths, but there was no intent to ask the committee to apply any stopping rules, and no reason to believe the interventions would cause harm (23). For each 24-h recall interview, respondents gave verbal consent and were free to stop the interviews at any point.

Statistical methods
Sample size. The target sample size was 200 households per arm, to detect a difference in the primary outcome (RDEAR) of 0.1 from 0.9 to 1.0 between 2 trial arms, with 80% power and 95% confidence. This assumed 19 clusters (allowing loss of 1 cluster) per arm, an SD of 0.27, 1.0 between 2 trial arms, with 80% power and 95% confidence. This assumed 19 clusters (allowing loss of 1 cluster) per arm, an SD of 0.27, and 95% confidence. This assumed 19 clusters (allowing loss of 1 cluster) per arm, an SD of 0.27, and 95% confidence. This assumed 19 clusters (allowing loss of 1 cluster) per arm, an SD of 0.27, and 95% confidence. This assumed 19 clusters (allowing loss of 1 cluster) per arm, an SD of 0.27, and 95% confidence.

Main analysis methods. Main analyses were performed under the intention-to-treat principle, so included all sampled households irrespective of their exposure to the interventions. Effect sizes (regression coefficient or OR) were estimated by fitting linear or logistic mixed-effects regression models. We treated clusters as a random effect and strata as a fixed effect. Allocation ratios were log-transformed to give normally distributed residuals. For log-transformed outcomes, effect sizes of transformed outcomes are reported in the tables, but the exponents are provided in the text for interpretability. Effect sizes, 95% CIs, and P values are reported. Cluster-adjusted chi-square tests assessed differences in characteristics between respondents and nonrespondents. We included outliers (>3500 or <1000 kcal) in energy intakes and data from respondents who were fasting or feasting; results were very similar when outliers were excluded. For MPA ratios only, we had some extreme outliers in log-transformed MPA ratios (values < −8), which we excluded to give normally distributed residuals.
TABLE 2  Characteristics of respondents and nonrespondents

|                   | Total eligible at start of study, % | Became ineligible during study | Eligible but not sampled | Sampled in all arms | Control | PLA | PLA + cash | PLA + food |
|-------------------|-------------------------------------|--------------------------------|--------------------------|---------------------|---------|-----|-----------|-----------|
| n                 | 2026                                | 952                            | 269                      | 805                 | 150     | 154 | 283       | 218       |
| Age of pregnant woman (n = 2015) |                                     |                                |                          |                     |         |     |           |           |
| <19 y             | 11.5                                | 12.9                           | 7.5                      | 11.2                | 12.7    | 14.9 | 10.6      | 8.3       |
| 19–29 y           | 81.9                                | 79.1                           | 82.7                     | 84.8                | 84.0    | 83.1 | 83.8      | 88.1      |
| ≥29 y             | 6.6                                 | 7.9                            | 9.8                      | 4.0                 | 3.3     | 2.0  | 5.7       | 3.7       |
| Gravity (n = 1945) |                                     |                                |                          |                     |         |     |           |           |
| Primigravid       | 35.8                                | 38.4                           | 29.6                     | 35.0                | 34.5    | 31.9 | 33.8      | 29.2      |
| Multigravid       | 64.2                                | 61.6                           | 70.4                     | 65.0                | 65.5    | 68.1 | 66.2      | 70.8      |
| Caste group (n = 2015) |                                   |                                |                          |                     |         |     |           |           |
| Dalit and Muslim (disadvantaged groups) | 34.9                                | 37.7                           | 38.3                     | 30.4                | 35.3    | 32.5 | 29.3      | 27.1      |
| Janjati/other Terai castes | 42.2                                | 40.1                           | 42.1                     | 44.7                | 42.7    | 54.5 | 41.0      | 44.0      |
| High caste (Yadav, Brahmin) | 22.9                                | 22.1                           | 19.5                     | 24.8                | 22.0    | 13.0 | 29.7      | 28.9      |
| Principal component wealth score of 12 items (n = 1981) |                                     |                                |                          |                     |         |     |           |           |
| Lowest tertile    | 31.5                                | 33.0                           | 38.1                     | 27.8                | 36.5    | 39.9 | 27.4      | 34.4      |
| Middle tertile    | 33.1                                | 31.8                           | 30.4                     | 35.5                | 33.1    | 30.1 | 35.9      | 32.6      |
| Highest tertile   | 35.3                                | 35.2                           | 31.5                     | 36.8                | 30.4    | 30.1 | 36.7      | 33.0      |
| Land ownership (n = 1981) |                                   |                                |                          |                     |         |     |           |           |
| Landless          | 31.2                                | 31.5                           | 39.6                     | 28.0                | 34.5    | 34.6 | 21.7      | 27.1      |
| Owns land         | 68.9                                | 68.5                           | 60.4                     | 72.0                | 65.5    | 65.4 | 78.3      | 72.9      |
| Maternal education (n = 1982) |                                   |                                |                          |                     |         |     |           |           |
| Never went to school | 58.8                                | 61.6                           | 63.5                     | 54.0                | 56.1    | 55.6 | 52.3      | 53.7      |
| Primary to lower secondary | 22.8                                | 20.7                           | 22.7                     | 25.2                | 27.0    | 26.1 | 27.4      | 20.6      |
| Secondary and above | 18.4                                | 17.7                           | 13.8                     | 20.8                | 16.9    | 18.3 | 20.3      | 25.7      |
| Husband’s education (n = 1981) |                                   |                                |                          |                     |         |     |           |           |
| Never went to school | 41.9                                | 44.7                           | 46.3                     | 37.4                | 42.2    | 34.4 | 39.6      | 30.7      |
| Primary to lower secondary | 32.3                                | 30.9                           | 32.8                     | 33.6                | 29.9    | 34.4 | 33.9      | 36.7      |
| Secondary and above | 25.8                                | 24.4                           | 20.8                     | 29.0                | 27.9    | 31.1 | 26.4      | 32.6      |
| Overseas migration (n = 1941) |                                   |                                |                          |                     |         |     |           |           |
| Husband not living overseas | 80.9                                | 78.7                           | 81.6                     | 83.3                | 78.8    | 81.3 | 87.7      | 81.9      |
| Husband living overseas | 19.1                                | 21.3                           | 18.4                     | 16.7                | 21.2    | 18.7 | 12.3      | 18.1      |

1 PLA, Participatory Learning and Action.  
2 Response rates for characteristics of sampled households ranged between 98% (789 of 805 for overseas migration) and 100%.  
3 Cluster-adjusted χ² test comparing respondents and non-respondents = 8.80; P = 0.066.  
4 Cluster-adjusted χ² test comparing respondents and non-respondents = 4.47; P = 0.107.  
5 Cluster-adjusted χ² test comparing respondents and non-respondents = 3.14; P = 0.534.  
6 Cluster-adjusted χ² test comparing respondents and non-respondents = 3.19; P = 0.527.  
7 Cluster-adjusted χ² test comparing respondents and non-respondents = 4.35; P = 0.114.  
8 Cluster-adjusted χ² test comparing respondents and non-respondents = 7.71; P = 0.107.  
9 Cluster-adjusted χ² test comparing respondents and non-respondents = 5.39; P = 0.250.  
10 Cluster-adjusted χ² test comparing respondents and non-respondents = 2.35; P = 0.107.

Covariate adjustment. To select covariate adjustments, goodness-of-fit for nested models was assessed with the use of Wald tests, comparing models with and without the covariate and excluding covariates in the model using a backward-stepwise method. Covariates included in the models were as follows: caste-religion group (3 categories), tertiles of a wealth score (the first component from a principal components analysis based on ownership of assets), maternal education, husband living overseas, and a binary variable indicating whether the interview was conducted before or during the monsoon season. Log-RDEAR was also included as a covariate in analyses for iron, MPA, and food item allocation. The following covariates were tested but excluded due to a lack of significant association with the outcomes: maternal age, gestational age, and household size. Adjusted analyses had a lower number of observations; 16 of 805 households were dropped due to missing data on overseas migration.

Dose-response. A dose-response analysis tested the effect of high exposure compared with lower exposure for outcomes from those primary analyses (RDEARs) showing significant effects. We compared the effect of pregnant women receiving ≥6 transfers with <6 transfers and any household member attending the group at least once with no one ever attending the groups.

The significance level for all hypothesis testing was set at 0.05, with no formal adjustment for multiple comparisons. All of the analyses were conducted with Stata SE 14 (StataCorp LP).

Results
Response rate and respondent characteristics
Reasons for nonresponse are shown in Figure 2.
Table 3: Exposure to PLA groups by arm, and receipt of food and cash transfers

| Transfer exposure | Control | PLA | PLA + cash | PLA + food |
|-------------------|---------|-----|------------|------------|
| 0–3 transfers     | 2.5     | 7.4 | 60.9       | 31.7       |
| 4–5 transfers     | 37.9    | 60.9| 60.9       | 31.7       |
| 6–7 transfers     | 59.6    | 31.7| 31.7       | 31.7       |

1Response rates were 99% for women’s groups, 98% for cash transfers, and 93% for food transfers. PLA, Participatory Learning and Action.

Table 4: Effects of PLA groups, PLA with cash, and PLA with food, on nutrient intakes of pregnant women

| Outcome | Control | PLA | PLA + cash | PLA + food |
|---------|---------|-----|------------|------------|
| Kilocalories/d | 2239 ± 730 (2146) | -118 [-329, 94] | 0.275 | -80 [-1267, 117] | 0.427 | 86 [-114, 286] | 0.399 |
| Kilocalorie adequacy ratio/d | 0.88 ± 0.29 (0.84) | -0.07 [-0.15, 0.01] | 0.090 | -0.05 [-0.12, 0.03] | 0.242 | 0.04 [-0.03, 0.12] | 0.279 |
| Protein adequacy ratio/d | 1.32 ± 0.46 (1.30) | 0.01 [-0.02, 0.03] | 0.577 | 0.03 [-0.00, 0.06] | 0.069 | 0.00 [-0.03, 0.03] | 0.929 |
| Dietary iron adequacy ratio/d | 0.42 ± 0.15 (0.42) | 0.01 [-0.05, 0.07] | 0.723 | 0.06 [0.01, 0.11] | 0.015 | 0.07 [0.01, 0.12] | 0.012 |
| MPA | 0.37 ± 0.20 (0.36) | 0.05 [0.00, 0.09] | 0.044 | 0.04 [0.00, 0.09] | 0.037 | -0.01 [-0.05, 0.03] | 0.639 |
| MDD-W (score of 0 to 10) | 4.6 ± 1.2 (5.0) | 0.05 [-0.25, 0.36] | 0.725 | 0.35 [0.08, 0.63] | 0.012 | 0.14 [-0.14, 0.43] | 0.316 |
| MUAC, cm | 23.5 ± 2.1 (23.5) | 0.75 [0.28, 1.23] | 0.002 | 0.75 [0.33, 1.17] | -0.001 | 0.49 [0.05, 0.93] | 0.029 |

1Values for the control arm are means ± SDs (medians), and effect sizes are adjusted differences (95% CIs). PLA, Participatory Learning and Action.
2Mean ± SDs (medians) were calculated by using mean intakes of the 3 recalls, rather than “usual” intakes calculated from best linear unbiased predictors. n = 805.
3Multivariable linear regressions adjusted for clustering, strata, caste/religion, tertiles of a wealth score, maternal education, husband migrating overseas, monsoon season (vs. premonsoon). n = 789.
4Modeled with log-transformed outcome. Multivariable linear regressions adjusted for clustering, strata, caste/religion, tertiles of a wealth score, maternal education, husband migrating overseas, monsoon season (vs. premonsoon), and log-transformed kilocalorie adequacy ratio. n = 789.
TABLE 5  Effects of PLA groups, PLA with cash, and PLA with food on pregnant women’s intakes of key food types and iron-folate supplements

| Outcome                      | Control | OR relative to control (95% CI) | P       | PLA          | OR relative to control (95% CI) | P       | PLA + cash | OR relative to control (95% CI) | P       | PLA + food | OR relative to control (95% CI) | P       |
|------------------------------|---------|---------------------------------|---------|--------------|---------------------------------|---------|-------------|---------------------------------|---------|-------------|---------------------------------|---------|
| Iron-folate supplements      | 28.6    | 2.53 (1.14, 5.60)               | <0.02   | 4.62 (2.19, 9.78) | <0.001                          |         | 3.08 (1.45, 6.54) | 0.004                          |         |
| Fresh foods                  | 32.7    | 1.28 (0.78, 2.11)               | <0.33   | 1.53 (0.98, 2.40) | <0.062                          | 1.39 (0.87, 2.22) | 0.165 |         |                                        |         |
| Dairy                        | 68.0    | 0.65 (0.39, 1.10)               | <0.108  | 1.80 (1.09, 2.88) | <0.022                          | 0.79 (0.48, 1.30) | 0.353 |         |                                        |         |
| Green leafy vegetables       | 66.7    | 1.01 (0.55, 1.86)               | <0.974  | 1.11 (0.64, 1.84) | <0.713                          | 1.22 (0.68, 2.17) | 0.502 |         |                                        |         |

1 PLA: Participatory Learning and Action.
2 Values for control arm are percentages, calculated based on any consumption over the repeated dietary recalls. n = 805.
3 Effect sizes are adjusted ORs relative to control (95% CIs). Odds are based on any consumption over the repeated dietary recalls. Multivariable logit regression models adjusted for clustering, strata, caste/religion, tertiles of a wealth score, maternal education, husband migrating overseas, monsoon season (vs. premonsoon). n = 789.

Trial effects on dietary intakes of pregnant women

Table 4 shows effects of the interventions on pregnant women’s intakes of key nutrients and nutritional status (MUAC), and Table 5 shows effects on food item consumption. Supplemental Table 1 reports arm-wise intakes of energy, protein, and 11 micronutrients, MPA, dietary diversity, and MUAC. Supplemental Table 2 reports the percentage of pregnant women consuming any of the 10 food groups or iron-folate supplements in each arm.

Compared with the control, participants consumed, on average, 86 kcal more in the PLA plus food arm and 118 kcal less in the PLA-only arm, but neither these nor the slight differences in kilocalorie adequacy ratios were significant. Dietary iron adequacy was higher in both transfer interventions than in the control arm (PLA plus cash = coefficient of 0.06; 95% CI: 0.01, 0.11; P = 0.015; PLA plus food: 0.07; 95% CI: 0.01, 0.12; P = 0.012). Notably, in the PLA plus cash arm, odds of consuming iron-folate supplements were 4.6 times higher, the odds of consuming dairy were 1.8 times higher, and the mean dietary diversity and PA were higher than in the control arm. For all interventions, the pregnant women’s mean MUAC and the odds of consuming iron-folate supplements were significantly higher than in the control arm.

Trial effects on intrahousehold food allocation

The trial effects on intrahousehold nutrient allocation are given in Table 6 (with the nutrient allocation ratios in each arm shown in Supplemental Table 3), and effects on intrahousehold food item allocation are reported in Table 7 (with food allocation ratios in each arm given in Supplemental Table 4). Effect sizes are reported as log-transformed outcomes in the tables, but their exponents are reported in the text for ease of interpretation.

In the PLA plus food arm, RDEARs between pregnant women and their mothers-in-law were significantly higher (12% higher; P = 0.014) than in the control arm, and RDEARs between pregnant women and their household heads were slightly but not significantly higher (8% higher; P = 0.078). On the contrary, in the PLA-only arm, RDEARs between pregnant women and the male household heads were 10% lower (P = 0.020), but MPA ratios (after adjusting for RDEARs) were 27% higher (P = 0.038) compared with the control arm. In the PLA plus cash arm, there was no difference in RDEARs or MPA ratios between pregnant women and household heads or their mothers-in-law relative to control.

For total iron (including supplements), compared with the control arm, all intervention arms had higher allocation ratios between pregnant women and male household heads (51% higher for PLA plus cash (P < 0.001); 24% higher for PLA only (P = 0.032) and PLA plus food (P = 0.022)). A similar trend was observed between pregnant women and their mothers-in-law. There were no trial effects on intrahousehold dietary iron allocation.

In all of the intervention arms, the odds of pregnant women consuming more animal-source foods (flesh foods and dairy) than the household head were higher than in the control arm, with the largest effect in the PLA-only arm (flesh food OR: 2.4; 95% CI: 1.09, 5.10; P = 0.028). There were no differences between trial arms in the odds of pregnant women consuming more flesh foods, dairy, or green leafy vegetables compared with their mothers-in-law.

Dose-response

There were no significant effects of different levels of exposure within arms on RDEARs. All coefficients were nonsignificant (all P values >0.3; results not shown).

Discussion

PLA groups, with and without food and cash transfers, are broadly beneficial for maternal nutrition and equity of intrahousehold food allocation, in a context in which the allocation of foods and nutrients are found to be highly inequitable. All of the interventions increased the consumption of iron-folate supplements, anthropometric status, and the intrahousehold allocation of some animal-source foods. The greatest effects on the intrahousehold allocation of energy between pregnant women and their mothers-in-law were in the
| Outcome                              | Control | PLA + cash | PLA + food |
|-------------------------------------|---------|------------|------------|
| **Mean ± SD** (median)              |         |            |            |
| r in adjusted n ≠ 0.00              | 803     | 803        | 803        |
| Adjusted difference relative to control (95% CIs) | | | |
| P                                   | 0.020   | 0.021      | 0.021      |
| P                                   | 0.020   | 0.021      | 0.021      |

**TABLE 6**

The effects of PLA groups, PLA with cash, and PLA with food on intrahousehold nutrient allocation.

PLA plus food arm, and the food transfers were channeled to pregnant women. However, the consumption of the food transfers was less than one-third of that received, and pregnant women’s average energy intakes did not significantly increase. In the PLA plus cash arm we observed the greatest beneficial effects on dietary diversity and the consumption and allocation of iron and dairy foods. Micronutrient adequacy also improved slightly, although energy intakes did not improve, nor did the allocation of energy between pregnant women and other household members. PLA groups without transfers had mostly positive but mixed effects on diets and dietary allocation, with notable improvements in MPA, MPA allocation, and the allocation of flesh foods. However, PLA group attendance was low in this arm at this point in the trial, and we also found, surprisingly, a slightly less equitable allocation of energy between pregnant women and male household heads. Given that energy deficiency is likely to be a bigger predictor of low birth weight than micronutrient deficiencies, the effects on RDEARs may explain why the PLA plus food intervention improved birth weight but the other interventions did not (24). Particularly in the PLA plus cash arm, the interventions may have resulted in other improvements in micronutrient status that were not measured.

Our results are consistent with an international review that found that cash transfers increase food expenditures, whereas food transfers increase energy consumption (10), perhaps because people spend their cash on more expensive (possibly more micronutrient-rich) but less energy-dense foods than the foods provided in transfers. The agreement of our results with this international review suggests that the reported South Asian norms that may restrict the allocation of food and cash transfers (food sharing, control over cash) are similar in other contexts or that PLA groups changed these norms to become similar to other contexts.

The effects of our interventions on iron intake and allocation may be explained by increased demand for iron-folate supplements or increased supply by FCHVs, perhaps due to PLA messages. Women may have also spent their transfers on iron-rich foods or supplements if they did not get supplements for free from the FCHVs. The effects on dairy consumption and allocation in the PLA plus cash arm were expected because the PLA groups encouraged women to spend their cash transfers on milk. In addition to being nutrient-rich and affordable, women could purchase milk from door-to-door sellers, thereby overcoming barriers that women face in leaving their homes. Fruit consumption also appears to be higher in the PLA plus cash arm, mirroring the PLA messaging to consume fruit. The small impact on overall dietary diversity, but larger improvements in consumption of certain foods, also mirrors findings from a cash transfer program in Bangladesh that found no impact on child dietary diversity but larger consumption of protein-rich foods (47) and a cash transfer program from Pakistan that found minimally higher food consumption but large increases in meat and fruit consumption (48).

In the PLA plus food arm, the positive effect on RDEARs is consistent with the evidence that food transfers increase energy consumption (10). Other household members rarely consumed the Super Cereal, indicating that it was not given preferentially to the typically favored household heads. Although the Super Cereal may have been given to people who were not interviewed, the observed channeling to pregnant women reflects the messaging in the PLA groups: that the Super Cereal should be treated like a “women’s medicine” for pregnancy. Qualitative research indicates that the Super Cereal was also considered inferior than the traditionally favored rice
Interventions by the World Food Programme in the west and if rice transfers were provided, such as the food-for-assets distribution and consumption amounts may have been observed acceptable and commonly consumed in this context. Different (49), although the Super Cereal was used to make foods that are considered unpopular in this cultural setting. The impacts of consumption patterns may have been observed earlier in trimesters, seasons, or earlier in the intervention period. Super Cereal may also have substituted other foods. This is difficult to determine because the difference in energy intakes between arms was small, although the consumption of other food groups in the PLA plus food arm appears mostly similar or higher than the PLA-only arm. An alternative explanation for the positive effects on RDEARs is that, because RDEARs account for energy requirements, the interventions caused reductions in pregnant women's workloads and energy requirements, as this was encouraged in the group meetings. However, exploratory analyses found no effect of any interventions on self-reported activity levels.

TABLE 7 Effects of PLA groups, PLA with cash, and PLA with food on intrahousehold food allocation

| Outcome | Control<sup>2</sup> | PLA<sup>3</sup> | PLA + cash<sup>3</sup> | PLA + food<sup>3</sup> |
|---------|----------------|----------------|----------------|----------------|
| Pregnant women vs. household heads | % households where pregnant women ate more | OR relative to control (95% CI) | P | OR relative to control (95% CI) | P | OR relative to control (95% CI) | P |
| Flesh foods | 7.3 | 2.36 (1.09, 5.10) | 0.028 | 1.86 (0.90, 3.85) | 0.093 | 2.20 (1.06, 4.55) | 0.034 |
| Dairy foods | 27.3 | 1.07 (0.60, 1.93) | 0.610 | 1.71 (1.02, 2.85) | 0.041 | 1.12 (0.66, 1.89) | 0.564 |
| Green leafy vegetables | 32.0 | 1.50 (0.63, 1.75) | 0.844 | 1.26 (0.81, 1.96) | 0.304 | 1.00 (0.63, 1.59) | 0.995 |
| Pregnant women vs. mothers-in-law | % households | OR relative to control (95% CI) | P | OR relative to control (95% CI) | P | OR relative to control (95% CI) | P |
| Flesh foods | 14.7 | 1.47 (0.79, 2.73) | 0.227 | 1.47 (0.84, 2.57) | 0.174 | 1.47 (0.82, 2.62) | 0.193 |
| Dairy foods | 40.7 | 1.08 (0.63, 1.80) | 0.625 | 1.30 (0.82, 2.07) | 0.270 | 0.93 (0.57, 1.50) | 0.759 |
| Green leafy vegetables | 33.3 | 1.13 (0.68, 1.88) | 0.633 | 1.11 (0.71, 1.73) | 0.649 | 1.22 (0.77, 1.95) | 0.392 |

1 PLA, Participatory Learning and Action; RDEAR, Relative Dietary Energy Adequacy Ratio.
2 Values are percentages of households where pregnant women ate more, based on average consumption of ≤3 dietary recalls per person. n = 805.
3 Effect sizes are adjusted ORs relative to control (95% CIs), based on average consumption of ≤3 dietary recalls per person. Multivariable logit regression models adjusted for clustering, strata, caste/religion, tertiles of a wealth score, maternal education, husband migrating overseas, monsoon season (vs. premonsoon), and log-RDEAR. n = 787.

(49), although the Super Cereal was used to make foods that are acceptable and commonly consumed in this context. Different distribution and consumption amounts may have been observed if rice transfers were provided, such as the food-for-assets interventions by the World Food Programme in the west and far-west of Nepal (50). Indeed, a comparison of rice and wheat transfers from Bangladesh found that rice was preferentially given to men but wheat was channeled to women (9).

Despite this channeling toward pregnant women, it is not clear that their consumption of Super Cereal caused these changes in RDEARs. Overall, Super Cereal consumption was low, perhaps because the sugar added makes the flour taste different and it is harder to cook with than unfortified flour. In addition, women were sampled in their third trimesters and they may have felt too full to consume this energy-dense food at that stage of pregnancy. Different consumption patterns may have been observed in earlier trimesters, seasons, or earlier in the intervention period. Super Cereal may also have substituted other foods. This is difficult to determine because the difference in energy intakes between arms was small, although the consumption of other food groups in the PLA plus food arm appears mostly similar or higher than the PLA-only arm. An alternative explanation for the positive effects on RDEARs is that, because RDEARs account for energy requirements, the interventions caused reductions in pregnant women's workloads and energy requirements, as this was encouraged in the group meetings. However, exploratory analyses found no effect of any interventions on self-reported activity levels.

We cannot disentangle the effects of PLA groups and transfers within arms. Group attendance was higher in these arms because the transfers were delivered at the groups and so incentivized attendance. In addition, PLA groups may have influenced the allocation of transfers, by enabling or instructing women to spend the cash on themselves or to selectively consume the Super Cereal rather than contributing the transfers to a shared household pot (49).

The PLA groups without transfers had some effects on nutrition, including higher MUAC (0.8 cm), allocation of flesh foods to pregnant women, and odds of consuming of supplements. There may also have been some increases in egg consumption in all arms, which is surprising because, although eggs were promoted in the PLA meetings, they are considered unpopular in this cultural setting. The impacts of PLA groups are consistent with studies from Bangladesh and India that found slight, positive effects of PLA groups on diet diversity (21, 22). However, there were lower RDEARs between pregnant women and household heads in the PLA-only arm relative to control, which is surprising because groups were hypothesized to encourage better care and diets. The higher anthropometric status, perhaps caused by PLA groups in earlier trimesters, may have increased women's energy requirements and made adequacy and equity harder to achieve compared with the control arm. Nevertheless, the possibility that households became less equitable and that women consumed less in the PLA-only arm cannot be ignored.

Some study limitations must be considered. The response rate and sampling frame limit the external validity of our results. However, formative research indicates that the sampled households, in which the women were living with their in-laws, were the strictest and most inequitable (14). Therefore, the effect sizes may be a conservative estimate if extrapolated to the whole population. We were unable to disentangle the effects of different household members attending the groups. The high attendance by mothers-in-law could explain some of these dietary improvements for pregnant women, because mothers-in-law often dictate the allocation of foods. Alternatively, attendance by mothers-in-law may have posed a barrier to pregnant women making new friends and speaking up in the group meetings—key processes hypothesized to influence dietary change. Testing effects across multiple outcomes increases the risk of spuriously reporting significant results for individual outcomes, but the outcomes are not independent and we have interpreted the results holistically by considering all outcomes together. We initially planned to collect data over 1 y, which would have provided impact estimates that are representative of different seasons. Instead, the evaluation took place over 4 mo at the end of the trial, and enthusiasm toward the interventions might have waned; only 25% of women ever attended PLA groups, compared with 55% in the main trial. Larger effects might have been detected earlier, particularly for green leafy vegetables that were in season earlier in the year (51). On the other hand, evaluation at this point allowed time for the interventions to have worked at individual, household, and community levels. We did not apply nutrient retention factors, which could bias estimates if the interventions selectively increased consumption of raw or cooked foods, although we expect this potential bias...
to be minimal. The use of average recipes prevents us from detecting effects on recipe composition and portions, such as channeling the choicest pieces of meat. Finally, we only sampled 3 household members; the interventions may have differentially influenced food allocation between pregnant women and junior members of the household.

In summary, all of our PLA interventions (with or without transfers) improved diets and MUAC for pregnant women. The PLA with cash improved dietary diversity and adequacy, whereas PLA with food improved equity in energy allocation. PLA interventions with and without cash or supplements have an important role to play in marginalized populations where food security and nutrition for pregnant women are poor. Scale-up is a policy option, but more research is needed to understand and improve participants’ perceptions of and adherence to the interventions.

Acknowledgments

We thank Sarah Style, Jayne Harthan, and Lu Gram from University College London and thank Ritesh Shrestha, Bidur Thapa, Rishi Neupane, and Bishnu Bhandari for their valuable inputs and support with logistics, planning, and intervention implementation. Lalbabu Kushiwata and Dhruva Devkota of Save the Children helped us to implement the food and cash transfer system. Pramila Ghimire, Srijana Nakermi, Marco Cavalcante, and Allison Prather of the World Food Programme provided the Super Cereal for the study and advice on managing its distribution. We thank Saska DePee from World Food Programme and Andrew Hall from Save the Children for providing technical advice. We thank Elaine Ferguson and Sarah Kehoe who provided comments on an earlier draft. Special thanks to the supervisors, Devlal Bishwakarma, Shankar Kapble, Suresh Mandal, and Dharambir Pasman. The authors’ responsibilities were as follows—HAH-F: wrote the manuscript and had primary responsibility for final content; HAH-F, NMS, AMdLC, and DSM: designed the research; HAH-F, PP, SJ, NS, and BPS: conducted the research; BJB, NS, and TH: provided essential materials; HAH-F, NMS, AC, and MC-B: analyzed data; and all authors: read and approved the final manuscript.

References

1. Black RE, Allen LH, Bhutta ZA, Caulfield LE, De Onis M, Ezzati M, Mathers C, Rivera J. Maternal and child undernutrition: global and regional exposures and health consequences. Lancet 2008;371(9608):243–60.
2. Black RE, Victora CG, Walker SP, Bhutta ZA, Christian P, De Onis M, Ezzati M, Grantham-McGregor S, Katz J, Martorell R. Maternal and child undernutrition and overweight in low-income and middle-income countries. Lancet 2013;382(9890):427–51.
3. Bastagli F, Hagen-Zanker J, Harman L, Barca V, Sturge G, Schmidt T, Pellerano L. Cash transfers: what does the evidence say? A rigorous review of programme impact and the role of design and implementation features. London: Overseas Development Institute; 2016.
4. Torheim LE, Ferguson EL, Penrose K, Arimond M. Women in resource-poor settings are at risk of inadequate intakes of multiple micronutrients. J Nutr 2010;140(11 Suppl):2051S–8S.
5. Han Z, Mullas S, Beyene J, Liao G, McDonald SD. Maternal underweight and the risk of preterm birth and low birth weight: a systematic review and meta-analyses. Int J Epidemiol 2011;40(1):63–101.
6. Hidrobo M, Hoddinott J, Peterman A, Margolies A, Moreira V. Cash, food, or vouchers? Evidence from a randomized experiment in northern Ecuador. J Dev Econ 2014;107:144–56.
7. Imdad A, Bhutta ZA. Maternal nutrition and birth outcomes: effect of balanced protein-energy supplementation. Paediatr Perinat Epidemiol 2012;26(Suppl 1):178–90.
8. Barber SL, Gertler PJ. Empowering women: how Mexico's conditional cash transfer programme raised prenatal care quality and birth weight. J Dev Effect 2010;2(1):51–73.
9. Ahmed AU, Qusimbing AR, Hoddinott JF, Nasreen M, Bryan E. Relative efficacy of food and cash transfers in improving food security and livelihoods of the ultra-poor in Bangladesh. Washington (DC): International Food Policy Research Institute; 2007.
10. Gentilini U. Our daily bread: what is the evidence on comparing cash versus food transfers? Washington (DC): World Bank Group; 2014. (Social protection and labor discussion paper 1420.)
11. Harris-Fry H, Shrestha N, Costello A, Saville NM. Determinants of intra-household food allocation between adults in South Asia—a systematic review. Int J Equity Health 2017;16:107.
12. Berti PR. Intra-household distribution of food: a review of the literature and discussion of the implications for food fortification programs. Food Nutr Bull 2012;33(3):516–570.
13. Harris-Fry HA, Paudel P, Shrestha N, Harrissont T, Beard BJ, Jha S, Shrestha BP, Manandhar DS, Costello AM, Cortina-Borja M et al. Status and determinants of intra-household food allocation in rural Nepal. Eur J Clin Nutr 2018;1.
14. Morrison J, Dural S, Harris-Fry H, Basnet M, Sharma N, Shrestha B, Manandhar DS, Costello A, Osirin D, Saville N. Formative qualitative research to develop community-based interventions addressing low birth weight in the plains of Nepal. Public Health Nutr 2017;21(2):377–84.
15. Girard AW, Olude O. Nutrition education and counselling provided during pregnancy: effects on maternal, neonatal and child health outcomes. Paediatr Perinat Epidemiol 2012;26(Suppl 1):191–204.
16. Nair N, Tripathy P, Sachdev HS, Bhattacharyya S, Gope R, Gagari S, Rath S, Rath S, Sinha R, Roy SS. Participatory women's groups and counselling through home visits to improve child growth in rural eastern India: protocol for a cluster randomised controlled trial. BMC Public Health 2015;15(1):384.
17. Nguyen PH, Kim SS, Sanghvi T, Mahmud Z, Tran LM, Shabnam S, Akhtar B, Haque R, Afzana K, Frongillo EA. Integrating nutrition interventions into an existing maternal, neonatal, and child health program increased maternal dietary diversity, micronutrient intake, and exclusive breastfeeding practices in Bangladesh: results of a cluster-randomized program evaluation. J Nutr 2017;147(12):2326–37.
18. Manandhar DS, Osirin D, Shrestha BP, Mesko N, Morrison J, Tumbahangnge KM, Tamang S, Thapa S, Shrestha D, Thapa B et al.; MIRA Makwanpur Trial Team. Effect of a participatory intervention with women’s groups on birth outcomes in Nepal: cluster-randomised controlled trial. Lancet 2004;364:970–79.
19. Rath S, Nair N, Tripathy PK, Barnett S, Rath S, Mahapatra R, Gope R, Bajpai A, Sinha R, Costello A. Explaining the impact of a group’s led community mobilisation intervention on maternal and newborn health outcomes: the Ekjit trial process evaluation. BMC Int Health Hum Rights 2010;10(1):25.
20. Prost A, Colbourn T, Seward N, Azad K, Coomarasamy A, Copas A, Hough TA, Dawes L, Ellerbrok E, Kuddus A, Lewycka S. Women’s groups practising participatory learning and action to improve maternal and newborn health in low-resource settings: a systematic review and meta-analysis. Lancet 2013;381(9879):1736–46.
21. Harris-Fry HA, Azad K, Younes L, Kuddus A, Shaha S, Nahar T, Hossen M, Costello A, Fottrell E. Formative evaluation of a participatory women's group intervention to improve reproductive and women’s health outcomes in rural Bangladesh: a controlled before and after study. J Epidemiol Community Health 2016;7063–70.
22. Nair N, Tripathy P, Sahdev H, Pradhan H, Bhattacharyya S, Gope R, Gagari S, Rath S, Rath S, Sinha R. Effect of participatory women's groups and counselling through home visits on children's linear growth in rural eastern India (CARING trial): a cluster-randomised controlled trial. Lancet 2004;364:970–79.
23. Saville NM, Shrestha BP, Style S, Harris-Fry H, Beard BJ, Sengupta A, Jha S, Rai A, Paudel V, Pulikki-Brannstrom A-M. Protocol of the MIRA Makwanpur Trial Team. Effect of a participatory intervention with women’s groups on birth outcomes in Nepal: cluster-randomised controlled trial. Lancet 2013;381(9879):1736–46.
24. Sabiven N, Shrestha BP, Style S, Harris-Fry H, Beard BJ, Sen A, Jha S, Rai A, Paudel V, Sah R et al. Impact on birthweight and child growth
