Infant mortality among native-born children of immigrants in France, 2008–17: results from a socio-demographic panel survey

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Background: Within Europe, France stands out as a major country that lacks recent and reliable evidence on how infant mortality levels vary among the native-born children of immigrants compared with the native-born children of two parents born in France. Methods: We used a nationally representative socio-demographic panel consisting of 296,400 births and 980 infant deaths for the period 2008–17. Children of immigrants were defined as being born to at least one parent born abroad and their infant mortality was compared with that of children born to two parents born in France. We first calculated infant mortality rates per 1000 live births. Then, using multi-level logit models, we calculated odds ratios of infant mortality in a series of models adjusting progressively for parental origins (M1), core demographic factors (M2), father’s socio-professional category (M3) and area-level urbanicity and deprivation score (M4). Results: We documented a substantial amount of excess infant mortality among those children born to at least one parent from Eastern Europe, Northern Africa, Western Africa, Other Sub-Saharan Africa and the Americas, with variation among specific origin countries belonging to these groups. In most of these cases, the excess infant mortality levels persisted after adjusting for all individual-level and area-level factors. Conclusions: Our findings, which can directly inform national public health policy, reaffirm the persistence of longstanding inequality in infant mortality according to parental origins in France and add to a growing body of evidence documenting excess infant mortality among the children of immigrants in France.

Introduction

The origin composition of births in France has transformed in the last few decades. Births to at least one foreign-born parent now contribute a third of all annual births—one of the largest shares in all of Europe—and the majority of this share have a parent, or parents, with origins outside of the EU. Despite this growing diversity, France stands out as a European country that lacks up-to-date evidence on how infant mortality levels vary according to parental origins. This is perhaps due to a national citizenship model that negates using such criteria to classify people and treat them as distinct groups in order to promote equality—a principle enshrined in public health care services. Despite this, evidence from government reports suggests that health inequalities are readily apparent from birth among the native-born children of immigrants, notably in the lower survival chances of children with parents from Africa. However, these reports are now dated, and their aim was simply to describe differentials in infant mortality according to the origins of the parents.

Thus, the objectives of this study are to contribute up-to-date findings on ‘how’ and ‘why’ infant mortality varies among the native-born children of immigrants in France for a recent period, by describing initial mortality levels according to detailed parental origins and by assessing the explanatory role of a range of relevant socio-demographic predictors. We place our work within a European context in which elevated infant mortality in the children of parents with non-EU origins is common. We add to the evidence by providing recent findings for a major European country. We extend it by contributing one of the first studies to adopt a multi-level framework examining the role of proximal (parental) and distal (socio-geographic) factors in variation in infant mortality according to parental origins. The lack of explanatory power of proximal factors in poor immigrant birth outcomes has re-focused attention on the role of more distal factors, which can affect child health ‘directly’ through enabling parental access to resources and services and ‘indirectly’ through their effect on parental socio-economic position. Considering prior evidence, we expect to find elevated infant mortality among the native-born children of immigrants in France.

Methods

Data

We used the French Permanent Demographic Sample (EDP), France’s largest socio-demographic panel and representative 4% sample of the population. The EDP contains individual-level anonymized records on life events from civil registers, censuses and geographic data at a fine-grained level of administrative division in France—the commune—facilitating multi-level analysis. Eligibility for the sample is based upon date of birth, being born on one of 16 specific days in the calendar year. New members enter the sample by being born in, or moving, to France and exit the sample by death or emigration. For our study, we only need to link the birth and death files between 1 January 2008 and 31
December 2016. Thus, our analysis is representative of all births in France—including documented and undocumented immigrants—
and is not selective upon, say, requiring an address in order to be
able to complete a census form. The EDP is entirely socio-
demographic and does not contain individual-level health data of
any kind.

**Study parameters**

Eligibility for the study is based upon being born in mainland France
between 1 January 2008 and 31 December 2016. Mainland France
refers to the area of the French Republic that is geographically in
Europe and not overseas France (i.e. French Guiana, Guadeloupe,
Martinique, Mayotte and Réunion). All children born in overseas
France and all children born in mainland France to a parent—or
parents—born in overseas France, are excluded from the analyses.
The outcome variable is binary and indicates whether each infant
born alive survived to his or her first birthday. The variable was
derived from calculating the exact age at death, itself derived from
the date of death (from the death file) minus the date of birth (from
the birth file).

The exposure variable is parental region of origin. The native-
born children of immigrants were defined as children born in main-
land France to at least one parent born abroad. The reference group
were defined as children born in mainland France to two parents
born in mainland France. To categorize children according to their
parental origins, we grouped countries into eight regional categories
using the United Nations M.49 classification. For the 3600 cases
(4.4% of the native-born children of immigrants) where parents
were born in different regions, we assigned children according to
mother’s region of origin. The composition of these groups can be
found in online Supplementary table S3, which shows the main
countries within each of these regional groups.

Alongside the parental region of origin, we adjusted for several
other individual-level predictor variables that included the sex of the
child (male vs. female), the year of birth (categorized into single
years from 2008 to 2016), the age of the mother (categorized into
5-year bands ranging from 15–19 to 40–49), single vs. multiple birth
(1 vs. 2+), and father’s socio-professional category (SPC; grouped as
executive and intellectual professions, intermediate professions, of-
ce workers, manual workers, farmers and self-employed including
craftsmen, small business leaders and shopkeepers). All of the pre-
dictor variables, including the parental country of birth, were
derived directly from data from the birth certificate in the EDP.

At the commune-level, we adjusted for the French Deprivation
Index (FDI), a composite indicator of neighbourhood deprivation
based upon average household income, percentage high school gradu-
ates in the population aged ≥15 years, percentage of blue-collar
workers in the active population and the unemployment rate. The
index was originally designed to be representative for the whole
of France and to take urban-rural comparability issues into account.25,26
The index is readily divided into quintiles of least, least, middle, more
and most deprived. As a measure of urbanicity, we adjusted for the
size of the urban unit, a spatial measure defined by the National
Institute of Statistics and Economic Studies (INSEE). This measure
classifies areas into rural areas of less than 2000 residents and
urban areas of 2000–19,999, 20,000–199,999, 200,000–1,999,999, and
2,000,000+ (the Paris region) residents. In both cases, we merged
these data into the EDP externally, attaching a category to each birth
based upon the mother’s place of residence as registered on the birth
certificate for the middle of the observation period, 2013.

The choice of the final individual-level predictors was based upon
an initial exploratory data analysis of available variables in the EDP.
For example, the birth file also contained information on birth order
and mother’s SPC. We initially considered these variables as import-
ant predictors that might help to explain group differences. However,
birth order was missing for 40% of births and mother’s SPC for 30%
of births (with higher levels among immigrant groups); these births
also had higher infant mortality and so we decided not to use them in
the analysis. The inclusion of several other predictors was explored
and dismissed due to moderate levels of missingness (≥20%; birth
interval and parental living arrangements) or a lack of association with
infant mortality (parental civil status). At the area-level, we considered
the % of the commune: in poverty, low educated, unemployed and/or
foreign-born. However, we found the level of urbanicity and com-
une deprivation score to have the most consistent associations with
infant mortality.

Due to the high quality of the data and the completeness of birth
and death certificates, we only have a small number of excluded
cases. Of the 303,260 births potentially eligible for inclusion in the
analyses, we dropped five cases in which no country of birth was
specified on the birth certificate. For parental country of birth, we
dropped 6500 (2%) cases in which we did not have the information
to be able to categorize children into a specific region or country of
origin. Finally, we removed 76 (0.03%) cases in which age of the
mother could not be derived from the birth data. This left a con-
formed eligible sample of 296,379 births in 26,434 communes with
980 infant deaths, with an average of 11 individuals per commune.

**Statistical methods**

To conduct our statistical analyses, we fitted multi-level logit models
for binary response variables. In Stata 15.0, we fitted a series of
nested mixed-effects logit models using ‘xmlologit’ with infant morta-
tality as our outcome, with children (level i) nested within commu-
nes (level j). The parameter estimates from the models were
exponentiated and interpreted as odds ratios (ORs). In Model 1,
we adjusted for parental region of origin only with commune-level
effects. In Model 2, we added our vector of individual-level demo-
graphic predictors: sex of child, year of birth, age of mother and
single vs. multiple birth. In Model 3, we added father’s socio-
professional category. In Model 4, we added size of urban unit
and deprivation score.

We also conducted two sensitivity analyses. To ensure our find-
ings were robust to different ways of defining the native-born chil-
dren of immigrants, we redefined them according to the mother’s
country of birth only, father’s country of birth only and contingent
on both parents being foreign-born. We re-fitted the region models
and compared them with the main results. Second, we fitted a com-
plete case analysis, removing the 12% of cases with missing father’s
SPC and compared the ORs to the main results. The findings from
these sensitivity analyses are documented at the end of the results
section and are available online.

**Results**

To assess the quality of the data, we compared the number of live
births and sex ratio of births in the study period to official estimates
from INSEE. Supplementary table S1 confirms an expected sampling
rate in the EDP of between 4.1% and 4.5% per year. An average sex
ratio of 1.05 in favour of men is identical to INSEE estimates. In
Supplementary table S2, the calculated sample infant mortality rate
(IMR) of 3.3 indicates only a minor underestimation compared to
3.6 in the INSEE estimates (a ratio of 0.92). A sample infant mortal-
ity rate (IMR) of 3.31 (3.10–3.52) reflects an average of
infant mortality rate (IMR) of 3.31 (3.10–3.52). We observe
higher IMRs in all groups except Northern and Western Europe and
Southern Europe, notably Eastern Europe [IMR = 5.25 (3.67–7.81)],

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Table 1 Births, deaths, infant mortality rates and parental origin composition of births in France, 2008–17

| Parental region of origin | Parental origin composition of all births (%) | Parental origin composition of births to immigrants only (%) | Births | Deaths | Infant mortality rate per 1000 |
|---------------------------|-----------------------------------------------|-------------------------------------------------|--------|--------|--------------------------------|
| All births                | –                                             | –                                               | 296 379| 980    | 3.31 (3.10–3.52)               |
| Births to two parents born in France | 72.1                                         | –                                               | 213 766| 631    | 2.95 (2.73–3.19)               |
| Births to at least one foreign-born parent | 27.9                                         | –                                               | 82 613 | 349    | 4.23 (3.79–4.69)               |

Births to at least one parent from:
- Northern and Western Europe: 1.5
- Southern Europe: 2.1
- Eastern Europe: 1.5
- Northern Africa: 11.9
- Western Africa: 3.1
- Other Sub-Saharan Africa: 2.9
- The Americas: 1.1
- Asia and Oceania: 3.9

| Parental origin | Births to at least one parent from | Deaths | Infant mortality rate per 1000 |
|----------------|-----------------------------------|--------|-------------------------------|
| Northern and Western Europe | 1.5                  | 4497   | 2.89 (1.54–4.94)             |
| Southern Europe       | 7.4                  | 6147   | 1.79 (0.89–3.20)             |
| Eastern Europe        | 5.5                  | 4571   | 5.25 (3.67–7.81)             |
| Northern Africa       | 42.5                 | 35 109 | 3.96 (3.33–4.68)             |
| Western Africa        | 11.0                 | 9052   | 7.07 (5.45–9.03)             |
| Other Sub-Saharan Africa | 2.9                 | 8472   | 4.72 (3.37–6.43)             |
| The Americas           | 3.8                  | 3149   | 5.40 (3.15–8.64)             |
| Asia and Oceania       | 14.1                 | 11616  | 3.53 (2.53–4.79)             |

Source: Authors’ calculations based upon Permanent Demographic Sample (EDP), 2008–17.

Western Africa [IMR = 7.07 (5.45–9.03)], and Other Sub-Saharan Africa [IMR = 4.72 (3.37–6.43)].

Table 2 shows the distribution of selected predictors by region of origin. Main trends include the much higher concentration of births to immigrants in large urban areas, the more favourable socioeconomic profiles for Northern and Western Europe and Americas, and less favourable socioeconomic profiles of Northern and Western Africa.

Figure 1 present the results from the multi-level logit models. We only show the results from Model 1 (which adjusts for parental region of origin only) and Model 4 (which additionally includes all of our individual and area-level predictors). The complete regression tables for Models 1–4 can be found in the Supplementary table S4. In all models, we investigate 980 instances of infant mortality among 296 379 births.

In Model 1, relative to children born to two parents born in mainland France, we observe excess infant mortality among those born to at least one parent from Eastern Europe [OR = 1.76 (1.14–2.69)], Northern Africa [OR = 1.29 (1.06–1.58)], Western Africa [OR = 2.42 (1.82–3.22)], Other Africa [OR = 1.61 (1.15–2.26)] and the Americas [OR = 1.91 (1.16–3.16)]. The size of these excess risks remain similar in Model 2 (with the addition of sex, year, multiple birth and age of mother), are attenuated in Model 3 and Model 4 if the child was male [OR = 1.32 (1.01–1.74)] and in urban units of 200 000–1,999 999 people [OR = 1.96 (1.26–3.05)], Senegal [OR = 2.22 (1.32–3.71)] and the DR Congo [OR = 2.87 (1.10–5.48)].

For the sensitivity analyses, Supplementary table S7 shows that we continue to see elevated infant mortality in the same region of origin groups regardless of which definition we use for the native-born children of immigrants. Similarly, a complete case analysis concerning missingness in father’s SPC from Supplementary table S8 shows that we continue to observe elevated infant mortality in the same region of origin groups as shown in figure 1.

Conclusions
The objectives of this study were to describe initial infant mortality levels among native-born children of immigrants in France and to determine how these levels were affected by adjusting for a range of socio-demographic predictors. To achieve this, we developed a detailed analysis based around a recent period to address the shortcomings of previous research in France. We documented excess infant mortality among children born to at least one parent from Eastern Europe, Northern Africa, Western Africa, Other Sub-Saharan Africa and the Americas. In the country-specific analysis, we found excess infant mortality levels among countries within these regions, namely Romania (Eastern Europe), Algeria (Northern Africa), Guinea, Mali, Senegal (all Western Africa) and DR Congo (Other Sub-Saharan Africa). In most cases, the excess mortality persisted after having adjusted for individual and area-level predictors.

Our study has many strengths. First, the analysis was based on a large-scale, reliable and representative socio-demographic panel survey. Second, we defined children of immigrants using a stable and large-scale, reliable and representative socio-demographic panel survey. Third, we provided some of the most granular estimates to date in France or beyond, which included both parental region and country of origin. Fourth, we adopted a novel multi-level approach that considered the
role of both individual- and area-level socio-economic factors in variation in infant mortality.

There are also several weaknesses. The sample size limited our ability to investigate trends over time (even if infant mortality levels in France have remained stable over the past decade)\textsuperscript{3,7}, examine causes of death or different forms of infant mortality (e.g. neonatal). We were also limited by small sample sizes, resulting in estimates for smaller country groups (e.g. Italy, Russia and China) with 95% CIs that ranged from sizeable advantages to disadvantages; these ORs should be interpreted with caution. Third, we could not adjust for certain predictors, either due to high levels of missingness (e.g. mother’s SPC and birth order) or because we lacked information (e.g. on health-related variables such as duration of gestation). This lack of more detailed individual-level parental SES predictors (e.g. education level) also limited our ability to capture other aspects of social disadvantage and explain more of the persisting excess mortality.

This study reaffirms that little progress has been made in addressing inequality in infant mortality among the native-born children of immigrants in France during the same period in which significant progress has been made in tackling social inequality.\textsuperscript{4,6} Our findings are consistent with a decade-old study that found elevated infant mortality levels among the native-born children of mothers with Sub-Saharan African nationalities\textsuperscript{8} and a 30-year-old study that found elevated infant mortality among the native-born children of parents with Western African and Northern African nationalities.\textsuperscript{8} Beyond providing recent estimates for France, we have extended the national evidence by documenting the emergence of excess mortality in children with parents from new regions for the first time (e.g. Eastern Europe), by indicating which countries are driving these excesses, and by examining possible socio-demographic causes.

Our data checks found the EDP to be representative of the situation at the national-level. With this in mind, our work represents a valuable addition to the small evidence base on the recent mortality situation of the native-born children of immigrants in France.\textsuperscript{28,29} Our work can help to reignite debate about the lack of progress made in the reduction of this inequality and inform new evidence-based policy. The findings also raise questions about a national citizenship model that advises against the identification of different population sub-groups in France to ensure equal treatment for all. It could be reasoned that the identification of such groups is necessary in order to measure understand, and attempt to address such inequality.

These findings add to the body of European literature identifying elevated infant mortality levels among native-born children of immigrants from non-EU and especially African countries in Belgium,\textsuperscript{13,14} Denmark,\textsuperscript{7,9} Italy,\textsuperscript{15} the Netherlands,\textsuperscript{18,19,30} Spain,\textsuperscript{21} Sweden,\textsuperscript{31} Switzerland\textsuperscript{16} and the United Kingdom.\textsuperscript{32,20} We extended this evidence by contributing some of the most detailed and robust findings for a European country lacking recent estimates and adopting a multi-level approach to consider the role of the broader socio-geographic context. Despite the widespread nature of this issue across Europe, there seem to be no specific EU policies in place to address it. The most recent European Perinatal Health Report, for example, made no explicit mention of the children of immigrants.\textsuperscript{32}

Finally, we adopted a multi-level framework in order to consider the role of proximal (the father’s SPC) and distal factors (urbanicity and the deprivation score of the commune) in variation in infant mortality levels. The rationale behind the approach was that variables detailing the wider socio-geographic context might help to

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**Table 2** Distribution of selected background characteristics (%) by parental origin composition in France, 2008–17

| Background characteristics | Mainland France | Northern Europe | Southern Europe | Eastern Europe | Northern Africa | Western Africa | Other Sub-Saharan Africa | The Americas | Asia and Oceania |
|---------------------------|----------------|----------------|----------------|---------------|----------------|---------------|----------------------------|-------------|----------------|
| Births                    | 213 766        | 4497           | 6147           | 4571          | 35 109         | 9052          | 8472                        | 3149        | 11 616         |
| Multiple birth            |                |                |                |               |                |               |                             |             |               |
| Yes                       | 3.3            | 4.1            | 3.6            | 3.4           | 3.4            | 3.6           | 3.3                         | 3.4         | 2.5            |
| Age of mother             |                |                |                |               |                |               |                             |             |               |
| 15–19                     | 2.1            | 1.2            | 3.0            | 5.6           | 1.3            | 2.6           | 2.3                         | 1.6         | 1.5            |
| 20–24                     | 13.6           | 9.2            | 14.4           | 15.9          | 13.6           | 15.3          | 13.8                        | 9.5         | 16.4           |
| 25–29                     | 33.9           | 23.7           | 26.5           | 28.4          | 29.5           | 30.4          | 29.3                        | 24.5        | 30.6           |
| 30–34                     | 32.8           | 35.1           | 30.3           | 31.1          | 30.4           | 29.1          | 30.6                        | 34.6        | 30.6           |
| 35–39                     | 14.6           | 24.2           | 19.8           | 15.3          | 19.3           | 17.0          | 18.4                        | 22.5        | 16.8           |
| 40–49                     | 3.0            | 6.7            | 6.1            | 3.8           | 6.0            | 5.7           | 5.7                         | 7.4         | 4.1            |
| Father’s SPC              |                |                |                |               |                |               |                             |             |               |
| Exec and intellectual     | 12.6           | 21.1           | 11.5           | 15.3          | 8.2            | 9.2           | 10.0                        | 20.7        | 11.1           |
| Intermediate              | 23.2           | 24.1           | 15.0           | 14.3          | 14.0           | 14.8          | 17.6                        | 22.6        | 13.4           |
| Office workers            | 12.4           | 10.0           | 7.4            | 7.6           | 10.6           | 16.3          | 14.5                        | 10.7        | 10.8           |
| Manual workers            | 33.3           | 23.9           | 45.7           | 33.1          | 42.1           | 37.2          | 34.6                        | 27.7        | 41.2           |
| Farmers                   | 1.8            | 1.2            | 0.7            | 1.2           | 0.3            | 0.5           | 0.7                         | 0.6         | 0.4            |
| Self-employed             | 6.7            | 8.6            | 7.0            | 7.1           | 8.9            | 5.0           | 4.9                         | 6.1         | 9.0            |
| Missing                   | 10.0           | 11.1           | 12.7           | 21.4          | 16.0           | 17.1          | 17.7                        | 11.6        | 14.1           |
| Size of urban unit        |                |                |                |               |                |               |                             |             |               |
| Rural                     | 26.5           | 21.2           | 11.1           | 7.6           | 2.8            | 3.5           | 4.3                         | 8.7         | 4.3            |
| Urban, 2–1999             | 17.8           | 14.4           | 12.7           | 8.6           | 7.2            | 4.8           | 6.0                         | 8.5         | 9.5            |
| Urban, 20–1999            | 17.1           | 15.7           | 17.8           | 15.7          | 19.1           | 12.6          | 15.7                        | 11.7        | 19.7           |
| Urban, 200–1999           | 24.5           | 26.9           | 27.4           | 31.2          | 37.4           | 20.1          | 30.4                        | 23.5        | 26.0           |
| Paris area                | 14.1           | 21.9           | 31.1           | 37.0          | 33.5           | 59.0          | 43.5                        | 47.6        | 40.5           |
| Deprivation index         |                |                |                |               |                |               |                             |             |               |
| Least                     | 19.2           | 32.1           | 26.5           | 27.1          | 15.6           | 23.1          | 21.9                        | 36.5        | 21.2           |
| Less                      | 20.1           | 19.2           | 19.9           | 19.6          | 17.8           | 22.5          | 21.4                        | 20.9        | 17.3           |
| Middle                    | 19.9           | 17.7           | 19.4           | 20.9          | 18.6           | 16.8          | 19.2                        | 15.0        | 16.3           |
| More                      | 20.6           | 15.9           | 16.6           | 14.4          | 18.6           | 13.6          | 16.8                        | 13.0        | 16.7           |
| Most                      | 20.2           | 15.1           | 17.6           | 18.1          | 29.4           | 23.9          | 20.8                        | 14.6        | 28.5           |

Source: Authors’ calculations based upon Permanent Demographic Sample (EDP), 2008–17.
capture broader aspects of social disadvantage (and more of the relative mortality excess) that would not be captured by the father’s SPC. While this proved to be the case, large amounts of excess remained in many of the groups. Consequently, we highlight several possible explanations (beyond the specified limitations relating to individual-level SES predictors) that could help to explain the persisting excesses.

In the literature, interpretations relate to the health-seeking behaviours among immigrants and to health-system related factors. Recent work found that Sub-Saharan African and Northern African mothers were at a greater risks of being overweight or obese prior to pregnancy and having gestational diabetes compared with mothers born in France, factors linked with increased infant mortality. In Norway and in the Netherlands, non-western immigrants
have been found to be less likely to attend antenatal care, leading to poorer detection of complications. Language barriers may also hamper the interpretation of clinical symptoms and lead to sub-optimal perinatal care, such as inadequate medication and refusal of Caesarean-sections. This has led some to conclude that, although antenatal care is free in many European countries, some sub-groups might face cultural and lingual barriers to utilizing it effectively.

Additionally, in France, discrimination in the workplace has been found to impact physical and mental health; discrimination has recently been linked with an elevated risk of preterm birth among women in Germany. For maternal and child health, Sub-Saharan
African women were found to have higher risks of stillbirth, preterm births and low birthweights than French women. Similarly, the risk of post-partum maternal death was twice as high among foreign women and highest for women from sub-Saharan Africa. These studies provided evidence that, among women who died, the level of care was more often considered sub-optimal for women with foreign nationality than for women with French nationality.

Further research is needed to understand what specific factors account for these excess infant mortality levels among the native-born children of immigrants in France. Special attention should be given to cultural factors influencing help-seeking behaviours, interactions with the health care system and communication barriers, sub-optimal health care and discrimination. Perinatal audits and detailed analyses of causes of deaths may also provide leads and new avenues for decision makers and public policies.

**Supplementary data**

Supplementary data are available at EURPUB online.

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**Key points**

- Children born in France to at least one parent born in Northern Africa, Western Africa, Other Sub-Saharan Africa, Eastern Europe and the Americas have elevated infant mortality levels.
- These elevated infant mortality levels persist after adjusting for a wide range of individual-level socio-demographic factors and area-level socio-economic and geographical factors.
- At an even more granular level, we find persistent excess infant mortality levels for specific countries in these regional groups, notably Algeria, Mali, Senegal, Guinea, DR Congo and Romania.
- These findings can directly inform public health policy in France, a country that lacks up-to-date and reliable estimates of this nature and enrich wider debates about the health status of immigrants and their children in Europe.

**References**

1 Les naissances en 2018 - Tableaux de séries longues – Les naissances en 2018 | Insee [Internet]. Available at: https://insee.fr/fr/statistiques/4190308?sommaire=4190525 (12 June 2020, date last accessed).
2 Naissances selon la nationalité et le pays de naissance des parents | Insee [Internet]. Available at: https://www.insee.fr/fr/statistiques/2381382 (12 June 2020, date last accessed).
3 Oberti M, The French republican model of integration: the theory of cohesion and the practice of exclusion. New Dir Youth Dev 2008;2006:55–74.
4 Nié X. Les Facteurs ExPLICATIFS de la Mortalité Infantile en France et Leur Évolution Récente/Explanatory Factors for Infant Mortality in France and Their Recent Development. Paris: Institut national de la statistique et des études économiques, 2011.
5 Barbieri M, Toulemon L. Les enfants tout égaux devant la mort? Problèmes d’observation et de mesure des différences sociales de la mortalité infantile en France? Are all children equal before death? Problems observing and measuring social differences in infant mortality in France. In: Histoires de Familles, Histoires Familiales Les Résultats de L’enquête Famille de Ined, 156è edn. Paris: Institut national d’études démographiques, 2005: 407–20.
6 Quang Chi D. Les inégalités sociales de la mortalité infantile s’estompent/Social inequalities in infant mortality are fading. Estat 1998;31:49–106.
7 Villadsen SF, Mortensen LH, Andersen AMN. Ethnic disparity in stillbirth and infant mortality in Denmark 1981–2003. J Epidemiol Community Health 2009 Feb 163:106–12.
8 Villadsen SF, Sievers E, Andersen A-MN, et al. Cross-country variation in stillbirth and neonatal mortality in offspring of Turkish migrants in northern Europe. Eur J Public Health 2010;20:530–5.
9 Brehm Christensen M, Fredsted Villadsen S, Weber T, et al. Higher rate of serious perinatal events in non-Western women in Denmark. Dan Med J 2016;63:A5197.
10 Gillet E, Saens B, Martens G, Cammu H. Fetal and infant health outcomes among immigrant mothers in Flanders, Belgium. Int J Gynecol Obstet 2014;124:128–33.
11 Kinge JM, Kornstad T. Assimilation effects on infant mortality among immigrants in Norway: does maternal source country matter? Demogr Res 2014;31:779–812.
12 Quary G. Child and infant mortality in England and Wales - Office for National Statistics [Internet]. 2020. Available at: https://www.ons.gov.uk/peoplepopulationandcommunity/birthsdeathsandmarriages/births/bulletins/childhoodinfantafterbirths/2010 (12 June 2020, date last accessed).
13 Racape J, De Spiegelacere M, Alexander S, et al. High perinatal mortality rate among immigrants in Brussels. Eur J Public Health 2010;20:536–42.
14 Racape J, Schoenborn C, Sow M, et al. Are all immigrant mothers really at risk of low birth weight and perinatal mortality? The crucial role of socio-economic status. BMC Pregnancy Childbirth 2016;16:75.
15 Simeoni S, Frova L, De Curtis M. Inequalities in infant mortality in Italy. Ital J Pediatr 2019;45:11.
16 Wanner P, Bollini P. The contribution of the foreign population to the high level of infant mortality in Switzerland: a demographic analysis. BMC Pregnancy Childbirth 2017;17:151.
17 Gissler M, Alexander S, Macfarlane A, et al. Stillbirths and infant deaths among migrants in industrialized countries. Acta Obstet Gynecol Scand 2009;88:134.
18 Troe E-J, Bos V, Deerenberg JM, et al. Ethnic differences in total and cause-specific infant mortality in the Netherlands. Paediatr Perinat Epidemiol 2006;20:140–7.
19 Troe E-J, Kunst AE, Bos V, et al. The effect of age at immigration and generational status of the mother on infant mortality in ethnic minority populations in The Netherlands. Eur J Public Health 2007;17:134–8.
20 Bakeo AC. Investigating variations in infant mortality in England and Wales by mother’s country of birth, 1983–2001. Paediatr Perinat Epidemiol 2006;20:127–39.
21 Luque-Fernandez MA, Franco M, Gelaye B, et al. Unemployment and stillbirth risk among foreign-born and Spanish pregnant women in Spain, 2007–2010. Eur J Epidemiol 2013;28:991–9.
22 Kim D, Saada A. The social determinants of infant mortality and birth outcomes in western developed nations: a cross-country systematic review. Int J Environ Res Public Health 2013;10:2296–335.
23 Mohamoud YA, Kirby RS, Ehrenthal DB. Poverty, urban-rural classification and term infant mortality: a population-based multilevel analysis. BMC Pregnancy Childbirth 2019;19:40.
24 Kane JB, Miles G, Youkavich J, King K. Neighborhood context and birth outcomes: going beyond neighborhood disadvantage, incorporating affluence. SSM Popul Health 2017;3:699–712.
25 Winderberfer F, Rican S, Jougla E, Rey G. Spatialtemporal association between deprivation and mortality: trends in France during the nineties. Eur J Public Health 2012;22:347–53.
26 Rey G, Jougla E, Fouillet A, Hémon D. Ecological association between a deprivation index and mortality in France over the period 1997–2001: variations with spatial scale, degree of urbanicity, age, gender and cause of death. BMC Public Health 2009;9:33.
27 Papon S. La mortalité infantile est stable depuis dix ans après des décennies de baisse—Insee Focus—117 [Internet]. 2018. Available at: https://www.insee.fr/fr/statistiques/3560308 (17 June 2020, date last accessed).
28 Guillot M, Khlat M, Wallace M. Adult mortality among second-generation immigrants in France: results from a nationally representative record linkage study. Demogr Res 2019;40:1603–44.
The association between migration and smoke-free families: how do migrants from different world regions compare?

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Background: Studies on adolescent secondhand smoke exposure within the family often dichotomously operationalize migration background without paying attention to social and cultural diversity within migrant populations. As a result, little is known about variation within migrant groups in smoke-free family environments (SFFEs). This study analyses the association between SFFEs and parental migration from different world regions.

Methods: Data from 14- to 16-year-old adolescents (N = 17 144) on SFFEs and parental migration were obtained from cross-sectional repeated SILNE-R surveys. A multivariable multinomial regression was applied, presenting relative risks (RRs) with 95% confidence intervals (CIs) for maternal or paternal tobacco smoking and home smoking bans. Variation in migration background was measured according to parental sex and place of birth.

Results: Approximately 18% of adolescents are exposed to maternal smoking, and 25% are exposed to paternal smoking. Almost half of the respondents do not live in SFFEs but are subject to permissive (5%) or partial (39%) smoking bans at home. We found that adolescents of Eastern European descent are at a higher risk of being exposed to both paternal and maternal smoking. A sex difference in parental smoking was found among Arabic/Islamic migrants, where mothers are less likely to be smokers. Maternal and paternal African origins are associated with prohibitive smoking bans at home. Eastern European mothers show higher odds of permissiveness and freely allowing smoking at home.

Conclusion: Notable within-differences according to parental sex and place of birth were found for SFFEs and should be taken into account when implementing equity-sensitive tobacco prevention programs.

Introduction

According to the Global Youth Tobacco Surveys, conducted in 168 countries, 30% of never-smoking adolescents are exposed to secondhand smoke (SHS) in their homes. A total smoking ban at home and a fully smoke-free (SF) family environment in advanced tobacco control settings such as Finland is reported by 58%.