RESEARCH ARTICLE

Household overcrowding and risk of SARS-CoV-2: analysis of the Virus Watch prospective community cohort study in England and Wales [version 1; peer review: 1 approved, 2 approved with reservations]

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Abstract

Background: Household overcrowding is associated with increased risk of infectious diseases across contexts and countries. Limited data exist linking household overcrowding and risk of COVID-19. We used data collected from the Virus Watch cohort to examine the association between overcrowded households and SARS-CoV-2.

Methods: The Virus Watch study is a household community cohort of acute respiratory infections in England and Wales. We calculated overcrowding using the measure of persons per room for each household. We considered two primary outcomes: PCR-confirmed positive SARS-CoV-2 antigen tests and laboratory-confirmed SARS-
CoV-2 antibodies. We used mixed-effects logistic regression models that accounted for household structure to estimate the association between household overcrowding and SARS-CoV-2 infection.

**Results:** 26,367 participants were included in our analyses. The proportion of participants with a positive SARS-CoV-2 PCR result was highest in the overcrowded group (9.0%; 99/1,100) and lowest in the under-occupied group (4.2%; 980/23,196). In a mixed-effects logistic regression model, we found strong evidence of an increased odds of a positive PCR SARS-CoV-2 antigen result (odds ratio 2.45; 95% CI:1.43–4.19; p-value=0.001) and increased odds of a positive SARS-CoV-2 antibody result in individuals living in overcrowded houses (3.32; 95% CI:1.54–7.15; p-value<0.001) compared with people living in under-occupied houses.

**Conclusion:** Public health interventions to prevent and stop the spread of SARS-CoV-2 should consider the risk of infection for people living in overcrowded households and pay greater attention to reducing household transmission.

**Keywords**
COVID-19, Overcrowding, SARS-CoV-2
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Competing interests: ACH serves on the UK New and Emerging Respiratory Virus Threats Advisory Group. AMJ was a Governor of Wellcome Trust from 2011-18 and is Chair of the Committee for Strategic Coordination for Health of the Public Research. SVK is a member of the UK Government's Scientific Advisory Group on Emergencies (SAGE) subgroup on ethnicity and co-chair of the Scottish Government's Expert Reference Group on ethnicity and COVID-19.

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Introduction
Household overcrowding is associated with increased risk of infectious diseases across cultures and countries. The World Health Organization Housing and Health Guidelines emphasise the health risks of overcrowding and note its complex economic, social and political determinants. Several definitions of overcrowding exist and it is also used as an indicator of material deprivation. Measures of overcrowding assess whether there is adequate dwelling space for occupants' needs related to shelter, space and privacy.

According to the English Housing Survey, approximately 787,000 (3%) of English households are overcrowded with unequal distribution across regions and social groups. A total of 7% of the most deprived households were overcrowded compared with less than half a per cent of the least deprived households. Overcrowding was highest in London compared with all other English regions. White British households are less likely to be overcrowded than households from all other ethnic groups.

There is strong evidence that household overcrowding is associated with risk of infectious diseases, such as tuberculosis, and increasing evidence on the association between household overcrowding and coronavirus disease 2019 (COVID-19). Studies from Europe have not found a consistent association: Swedish registry data revealed increased COVID-19 mortality among over 70 year olds from households, or care homes, with fewer square metres per inhabitant, but a retrospective cohort in Spain did not. In the USA, an ecological analysis found that COVID-19 death rates were higher in the counties with highest percentage of household crowding (16.8 per 100,000) compared with the least crowded areas (4.9 per 100,000). A study using 2011 UK census data examined ethnicity, household composition and COVID-19 mortality and found that elderly adults living with younger people were at increased risk of COVID-19 mortality. The study described the distribution of overcrowded housing within census data used, but it did not adjust for overcrowded housing in the causal mediation analysis as these were considered consequences of living in a multi-generational household rather than confounding factors. A study of UK Biobank data found household size was associated with COVID-19 infection after adjusting for age, sex, deprivation, ethnicity and body mass index (BMI). The REACT study found severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2) antibodies increased from 4.7% for people in single-person households to 13% in households of seven or more. The Open-SAFELY study found household size accounted for 10–16% of the excess risk of testing positive for SARS-CoV-2 and 12–39% of the excess risk of COVID-19 mortality in South Asian groups.

Relatively few studies have investigated the impact of overcrowding specifically, as opposed to household size. In Birmingham (UK), an analysis of data from 408 hospitalised COVID-19 patients found that people from areas of the city with low housing quality and overcrowding were more likely to be admitted to intensive care compared with other areas. A study of ≥70 year olds using Swedish cause of death, administrative and dwelling registry data corroborates this finding with individual-level overcrowding data. After adjusting for individual age, sex, country of birth, income and education, COVID-19 mortality was 2.1-times (HR 2.1, 95% CI 1.53–2.87) higher among ≥70 year olds in households with less than 20 m²/inhabitant compared with those with >60 m²/inhabitant. In a serological survey undertaken in Lima, Peru, lateral-flow SARS-CoV-2 seropositivity was also more prevalent among households in the two most overcrowded quartiles compared with the least overcrowded quartile (respective prevalence ratios 1.41 and 1.99, 95% CI 1.01–1.97 and 1.41–2.81) as measured by a ratio of inhabitants to habitable rooms with adjustment for sex, age group, local region and socio-economic status.

Overcrowding, as defined by more than three household members with fewer than six rooms between them, carried increased odds of laboratory-confirmed secondary infection of SARS-CoV-2 in 92 North Carolinian households within 28 days of a PCR-confirmed index case. Higher rates of household overcrowding observed among non-White/Hispanic people were also argued to account for the higher rates of secondary transmission found among those ethnic groups in this study.

Additional evidence on the association between household overcrowding and COVID-19 risk can guide immediate public health mitigation measures to reduce the spread of the SARS-CoV-2 and inform long-term housing policy in the UK. In this analysis we aim to examine the association between overcrowded households and either polymerase chain reaction (PCR) confirmed SARS-CoV-2 or antibodies acquired through SARS-CoV-2 infection.

Methods
The Virus Watch study is a household community cohort of acute respiratory infections in England and Wales that started recruitment in June 2020. As of 28th February 2021, there were 46,937 participants in Virus Watch. A detailed description of recruitment methods has been described previously, but in summary, to recruit our sample we used a range of methods. We used the Royal Mail Post Office Address File to generate a random list of residential address lists that were sent recruitment postcards (n=3,914), we placed social media adverts on Facebook and Twitter (n=18,594), sent SMS messages (n=11,151) and letters to participants from their General Practitioners (n=3,803). Participants were followed up weekly by email with a link to an illness survey which asks about the presence or absence of symptoms that could indicate COVID-19 disease including respiratory, gastrointestinal and general infection symptoms. The weekly survey is also used to capture SARS-CoV-2 test results received from outside the study (e.g. via the UK test, trace and isolate system).

Laboratory cohort
Nested within this larger study is a sub-cohort of 10,330 adults (aged over 18 years) participating in monthly antibody testing who completed at-home capillary blood sampling kits sent via post on a monthly basis, and provided self-reported
vaccination data on a weekly basis, in addition to demographic and clinical data at baseline. Individuals were included in this analysis if they underwent antibody testing between 1st February 2021 and 28th August 2021 and completed the February 2021 monthly survey.

Monthly survey
The Virus Watch monthly survey includes demographic, psychosocial/behavioral, environmental and health-related questions beyond the scope of the weekly survey. Data used in the analysis in this manuscript are taken from the Virus Watch third monthly survey that was sent to participants on 9th February 2021 and occurred during the third national lockdown for both England and Wales. In this survey, we included a series of questions about participants’ housing status adapted from housing-related items in the 2011 England and Wales Census as this the most comprehensive study in terms of coverage of the housing characteristics collected. The survey items collected comprised of: accommodation type; whether accommodation was self-contained; number of rooms available for exclusive use by the household, excluding bathrooms, toilets, halls or landings, or rooms that can be used only for storage (e.g. cupboards); number of bedrooms (built or converted for use as bedrooms, even if not currently used as a bedroom); housing tenure and, if rented, details of the rental arrangement; and whether the accommodation had central heating. Additionally, we presented binary (yes/no) questions about whether participants’ accommodation had visible mould or fungus and/or damp spots on the walls or ceiling due to their association with respiratory tract infections\(^{48,19}\). Housing-related survey questions are presented in full in the supplementary appendix in the extended data.

Overcrowding measure
To investigate overcrowding, we first calculated the persons per room\(^ 4\) for each household, which is defined as the number of household occupants divided by the number of rooms, excluding kitchen or bathrooms. We chose persons per room for several reasons, including: it aligned with the data we collected in our monthly survey which was based upon the 2011 UK Census question items; and it is a widely used and valid measure of overcrowding\(^ 4\); and it has been found to have fair agreement with other measures of overcrowding\(^ 4\). We did not ask participants to exclude kitchens from their reporting of the number of rooms in their accommodation and therefore we have assumed that all accommodation had a kitchen and subtracted one from the total number of rooms reported by each household. As described in previous studies examining the validity persons per room as a measure of overcrowding\(^ 4\), we categorised it as: 1) under-occupied accommodation where the number of rooms was greater than the number of people; 2) balanced accommodation where the number of rooms was equal to the number of people; 3) overcrowded accommodation where the number of rooms was fewer than the number of people.

Outcomes
We considered two primary outcomes. First, we were able to determine PCR-confirmed SARS-CoV-2 disease in all participants tested by the national surveillance system in community settings from March 2020 to August 2021. Positive SARS-CoV-2 results were identified from linkage of patient demographic characteristics (name, date of birth, address, NHS number) to the national Second Generation Surveillance System for SARS-CoV-2. Second, we examined laboratory-confirmed SARS-CoV-2 antibodies acquired through infection, among a subset of participants who underwent antibody testing between February 2021 to August 2021. The main outcome variable was evidence of prior infection defined as a cut off index of 0.1 or more on the Roche Elecsys anti-N total immunoglobulin assay that measures seropositive for the Nucleocapsid protein.

Covariates
Our analysis strategy was informed by conceptual models (from BMJ and the UK government) that have previously described the possible pathways between ethnicity and socio-economic status, overcrowding and risk of SARS-CoV-2 infection and we developed a directed acyclic graph to inform covariate selection (Supplementary Appendix Figure S1 in the extended data). We used this and the rules outlined by VanderWeele as principles of confounder selection\(^ 18\). In our primary analysis, we considered age, sex, ethnicity, household income and geographical region as covariates (Model A). Age, sex, ethnicity and geographical region were derived from participants’ responses to demographic questions at study baseline. Household income was derived from the February 2021 monthly survey.

Statistical analyses
We undertook description of the characteristics of included participants. To model the association between the selected covariates and each outcome, we conducted univariable and multivariable analyses using mixed-effects logistic regression models with a household-level random effect to account for household-level variation not explained by the covariates using the glmer function in R \(^{4.0.3}\) (RRID:SCR_001905).

Sensitivity analyses
We have examined our assumption regarding kitchens in our calculation of overcrowding, by only excluding one room from households with two or more rooms in a sensitivity analysis. In addition to the primary mixed-effects model, we ran three other separate sensitivity analyses (all accounting for household-level clustering): first, a minimal sufficient adjustment set informed by our directed acyclic graph (Model B); second, a model with age, sex, households with children (a binary variable representing households with and without children), and number of close contacts outside the household as covariates (Model C); third, a model with age, sex, ethnicity, income and occupation as covariates (Model D).

Ethics
The Virus Watch study has been approved by the Hampstead NHS Health Research Authority Ethics Committee. Ethics approval number - 20/HRA/2320. Written informed consent for the collection and use of data was obtained from participants.
Role of the funding source
The study sponsor(s) had no role in study design; in the collection, analysis, and interpretation of data; in the writing of the report; and in the decision to submit the paper for publication. We confirm that all authors accept responsibility to submit for publication. For information security reasons, RWA, EF, ME, IC, VN, SB, TB, AA, WLFE, CG, PP, MS, AMDN and AH had full access to all individual level data in the study analyses and all other authors had access to aggregated data.

Results
On 9th February 2021, participants in the Virus Watch study were invited to take part in the monthly survey. We include responses to the survey for 26,367 participants (56.2%; 26,367/46,937) in our primary analyses. The median number of rooms per household was 6 (interquartile range 5–7) and the median number of householders was 2 (interquartile range 1–3). 4.2% of participants (1,100/26,367) were classified as living in overcrowded households and 7.9% (2,071/26,367) in balanced households. Participants responding to the monthly survey containing housing related questions were more likely to be over 65, of White British ethnicity and live alone or with one other than the general population (Table 1).

Between 8th March 2020 and 10th August 2021, 2,689 Virus Watch participants (5.7%; 2,689/46,937) had a positive SARS-CoV-2 antigen test through the national test and trace system. 1,229 participants with a PCR confirmed SARS-CoV-2 antigen

Table 1. Description of Virus Watch participants on 28th February 2021, comparing all Virus Watch participants to those who took part in the February 2021 survey and ONS census data for England and Wales.

| Characteristic      | All Virus Watch participants on 28th Feb 2021 | Virus Watch Participants completing monthly survey in Feb 2021 | ONS* (%) |
|---------------------|-----------------------------------------------|--------------------------------------------------------------|----------|
| All                 | 46,937 (100%)                                 | 26,364 (100%)                                               | (100%)   |
| Age group           |                                               |                                                              |          |
| 0–15                | 5,866 (12%)                                   | 2,356 (8.9%)                                                | (19.1%)  |
| 16–24               | 2,698 (5.7%)                                  | 1,104 (4.2%)                                                | (10.6%)  |
| 25–44               | 9,123 (19%)                                   | 3,511 (13%)                                                | (26.1%)  |
| 45–64               | 15,781 (34%)                                  | 9,365 (36%)                                                | (25.6%)  |
| 65+                 | 13,469 (29%)                                  | 10,028 (38%)                                               | (18.5%)  |
| Ethnicity           |                                               |                                                              |          |
| White British       | 33,533 (71%)                                  | 22,712 (86%)                                               | (80.5%)  |
| White Irish         | 585 (1.2%)                                    | 370 (1.4%)                                                 | (0.9%)   |
| White Other         | 2,117 (4.5%)                                  | 1,245 (4.7%)                                               | (4.4%)   |
| Mixed               | 776 (1.7%)                                    | 422 (1.6%)                                                 | (2.2%)   |
| South Asian         | 1,576 (3.4%)                                  | 597 (2.3%)                                                 | (5.3%)   |
| Other Asian         | 368 (0.8%)                                    | 191 (0.7%)                                                 | (2.2%)   |
| Black               | 256 (0.5%)                                    | 113 (0.4%)                                                 | (3.3%)   |
| Other Ethnicity     | 201 (0.4%)                                    | 106 (0.4%)                                                 | (1%)     |
| Prefer not to say   | 116 (0.2%)                                    | 47 (0.2%)                                                  | -        |
| Missing             | 7,409 (16%)                                   | 561 (2.1%)                                                 | -        |
| Sex**               |                                               |                                                              |          |
| Male                | 17,652 (38%)                                  | 11,733 (45%)                                               | (49.4%)  |
| Female              | 21,982 (47%)                                  | 14,138 (54%)                                               | (50.6%)  |
| Missing/Suppressed  | 7303 (16%)                                    | 493 (1.9%)                                                 | -        |
| Number of householders*** |                                             |                                                              |          |
| 1                   | 7,174 (15%)                                   | 4,672 (18%)                                                | (29.5%)  |
| 2                   | 20,821 (44%)                                  | 13,802 (52%)                                               | (34.5%)  |
test were included (4.9%; 1,229/26,367). The proportion of participants with a positive SARS-CoV-2 antigen result was highest in the overcrowded group (9.0%; 99/1,100) and lowest in the under-occupied group (4.2%; 980/23,196; Table 2).

In a mixed-effects logistic regression model that included age, sex, ethnicity, household income and geographical region as fixed effects and a household-level random effect we found strong evidence of an increased odds of having a positive SARS-CoV-2 antigen result in individuals living in overcrowded accommodation (odds ratio (OR): 2.45; 95% CI:1.43–4.19; p-value=0.001) and those in balanced homes (OR: 1.72; 95% CI: 1.13–2.63; p-value=0.012) compared with those in under-occupied homes (Table 3). The proportion of variation at the household level (i.e. intracluster correlation coefficient, or ICC) was 64.9%.

Between 26th February 2021 and 28th August 2021, we had SARS-CoV-2 anti-N total immunoglobulin assay antibody results for 10,662 (23%; 10,662/46,937) participants with housing

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| Characteristic | All Virus Watch participants on 28th Feb 2021 | Virus Watch Participants completing monthly survey in Feb 2021 | ONS* (%) |
|----------------|---------------------------------------------|-------------------------------------------------------------|----------|
| 3              | 7,047 (15%)                                 | 3,363 (13%)                                                | (15.4%)  |
| 4              | 8,146 (17%)                                 | 3,345 (13%)                                                | (13.9%)  |
| 5              | 2,783 (5.9%)                                 | 960 (3.6%)                                                 | (4.5%)   |
| 6              | 966 (2.1%)                                   | 222 (0.8%)                                                 | (2.1%)   |
| Region         |                                             |                                                             |          |
| East Midlands  | 3,690 (7.9%)                                 | 2,352 (8.9%)                                               | (4.5%)   |
| East of England| 9,308 (20%)                                  | 5,799 (22%)                                                | (12.4%)  |
| London         | 6,732 (14%)                                  | 3,390 (13%)                                                | (9.3%)   |
| North East     | 2,156 (4.6%)                                 | 1,278 (4.8%)                                               | (8.1%)   |
| North West     | 4,614 (9.8%)                                 | 2,809 (11%)                                                | (10%)    |
| South East     | 8,096 (17%)                                  | 5,049 (19%)                                                | (10.5%)  |
| South West     | 3,011 (6.4%)                                 | 1,945 (7.4%)                                               | (15.1%)  |
| Wales          | 1,042 (2.2%)                                 | 593 (2.2%)                                                 | (15.4%)  |
| West Midlands  | 2,318 (4.9%)                                 | 1,434 (5.4%)                                               | (9.5%)   |
| Yorkshire and The Humber | 2,056 (4.4%) | 1,275 (4.8%)                                               | (5.3%)   |
| Missing        | 3,914 (8.3%)                                 | 440 (1.7%)                                                 | -        |
| Household overcrowding category**** |                               |                                                             |          |
| Under occupied | 23,196 (88%)                                 | 23,195 (88%)                                               |          |
| Balanced       | 2,071 (7.9%)                                 | 2,070 (7.9%)                                               |          |
| Overcrowded    | 1,100 (4.2%)                                 | 1,099 (4.2%)                                               | (3%)     |
| Missing        | 20,570                                      |                                                             |          |

* ONS data for age and region drawn from Mid-2019 Estimates of the Population for the UK, England, and Wales, Scotland and Northern Ireland (figures for England and Wales).

** We suppressed some groups due to deductive disclosure in some groups.

*** ONS data for household size drawn from Families and Households in the UK 2019 (UK wide estimates).

**** ONS overcrowded households figures from https://www.ethnicity-facts-figures.service.gov.uk/housing/housing-conditions/overcrowded-households/latest

Note: the total number of participants in Table 1 and Table 2 differs because a small number of participants completed the Feb 2021 survey after March 2021.
Table 2. Description of SARS-CoV2 PCR (up to 10th August 2021) and antibody test results (up to 28th August 2021) for housing and demographic characteristics of Virus Watch participants.

|                        | SARS-CoV2 PCR | SARS-CoV2 antibody |
|------------------------|---------------|--------------------|
|                        | Negative / Not tested | Positive | Negative | Positive |
| All                    | 25,138 (95%) | 1,229 (4.9%) | 9,672 (91%) | 990 (9.3%) |
| Household overcrowding category |               |                     |           |           |
| Under occupied         | 22,216 (96%) | 980 (4.2%) | 8,995 (92%) | 830 (8.4%) |
| Balanced               | 1,921 (93%)  | 150 (7.2%) | 495 (82%)  | 105 (18%)  |
| Overcrowded            | 1,001 (91%)  | 99 (9.0%)  | 182 (77%)  | 55 (23%)   |
| Age Group              |               |                     |           |           |
| 0–15                   | 2,209 (94%)  | 147 (6.2%) | -         | -         |
| 16–24                  | 983 (89%)    | 121 (11%) | 167 (81%) | 39 (19%)  |
| 25–44                  | 3,243 (92%)  | 269 (8.7%) | 1,052 (84%) | 202 (16%) |
| 45–64                  | 8,898 (95%)  | 468 (5.0%) | 3,835 (89%) | 462 (11%) |
| 65+                    | 9,805 (98%)  | 224 (2.2%) | 4,618 (94%) | 287 (5.9%) |
| Sex                    |               |                     |           |           |
| Male                   | 11,220 (96%) | 514 (4.4%) | 4,151 (91%) | 421 (9.2%) |
| Female                 | 13,440 (95%) | 700 (5.0%) | 5,506 (91%) | 566 (9.3%) |
| Missing/suppressed     | 478 (97%)    | 15 (3.0%)  | 15 (83%)  | 3 (17%)   |
| Ethnicity              |               |                     |           |           |
| White British          | 21,666 (95%) | 1,049 (4.6%) | 8,859 (91%) | 843 (8.7%) |
| White Irish            | 357 (96%)    | 13 (3.5%)  | 127 (89%) | 16 (11%)  |
| White Other            | 1,190 (96%)  | 55 (4.4%)  | 386 (86%) | 63 (14%)  |
| Mixed                  | 391 (93%)    | 31 (7.3%)  | 71 (79%) | 19 (21%)  |
| South Asian            | 552 (92%)    | 45 (7.5%)  | 104 (75%) | 34 (25%)  |
| Other Asian            | 182 (95%)    | 9 (4.7%)   | 57 (92%) | 5 (8.1%)  |
| Black                  | 107 (95%)    | 6 (5.3%)   | 26 (84%) | 5 (16%)   |
| Other Ethnicity        | 101 (95%)    | 5 (4.7%)   | 31 (91%) | 3 (8.8%)  |
| Prefer not to say      | 45 (96%)     | 2 (4.3%)   | 11 (85%) | 2 (15%)   |
| Missing                | 547 (98%)    | 14 (2.5%)  | 0 (NA%)  | 0 (NA%)   |
| Household income combined for last 12 months | | | | |
| £0-9,999               | 847 (95%)    | 44 (4.9%)  | 321 (90%) | 36 (10%)  |
| £10,000-24,999         | 4,526 (96%)  | 206 (4.4%) | 2,059 (93%) | 154 (7.0%) |
| £25,000-49,999         | 7,275 (96%)  | 327 (4.3%) | 3,206 (91%) | 304 (8.7%) |
| £50,000-£74,999        | 3,970 (95%)  | 227 (5.4%) | 1,515 (88%) | 205 (12%) |
| £75,000-£99,999        | 2,017 (95%)  | 106 (5.0%) | 676 (88%) | 90 (12%)  |
| £100,000+              | 2,103 (95%)  | 122 (5.5%) | 666 (88%) | 88 (12%)  |
| Prefer not to say      | 1,336 (96%)  | 52 (3.7%)  | 571 (93%) | 41 (6.7%) |
| Characteristic                      | SARS-CoV2 PCR | SARS-CoV2 antibody |
|------------------------------------|---------------|--------------------|
|                                   | Negative / Not Tested | Positive | Negative | Positive |
| Missing                           | 3,064 (95%) | 145 (4.5%) | 658 (90%) | 72 (9.9%) |
| Geographical region               |               |                  |                     |
| East Midlands                      | 2,230 (95%) | 122 (5.2%) | 865 (92%) | 77 (8.2%) |
| East of England                    | 5,548 (96%) | 252 (4.3%) | 2,302 (92%) | 193 (7.7%) |
| London                            | 3,183 (94%) | 207 (6.1%) | 963 (81%) | 232 (19%) |
| North East                         | 1,209 (95%) | 69 (5.4%) | 494 (93%) | 38 (7.1%) |
| North West                         | 2,646 (94%) | 164 (5.8%) | 1,031 (89%) | 129 (11%) |
| South East                         | 4,843 (96%) | 206 (4.1%) | 1,998 (93%) | 153 (7.1%) |
| South West                         | 1,877 (97%) | 68 (3.5%) | 758 (95%) | 38 (4.8%) |
| Wales                              | 593 (100%) | 0 (0%) | 209 (93%) | 16 (7.1%) |
| West Midlands                      | 1,378 (96%) | 57 (4.0%) | 539 (91%) | 54 (9.1%) |
| Yorkshire and The Humber           | 1,194 (94%) | 81 (6.4%) | 506 (90%) | 59 (10%) |
| Missing                            | 437 (99%) | 3 (0.7%) | 7 (88%) | 1 (12%) |

Note: the total number of participants in Table 1 and Table 2 differs because a small number of participants completed the Feb 2021 survey after March 2021.
The proportion of participants with a positive SARS-CoV-2 antibody result was highest in the overcrowded group (23%; 55/237) and lowest in the under-occupied group (8.4%; 830/9,825; Table 2). In a mixed-effects logistic regression model that included age, sex, ethnicity, household income and geographical region as fixed effects and a household-level random effect, we found evidence of an increased odds of having a positive SARS-CoV-2 antibody result in individuals living in overcrowded houses (OR: 3.32; 95% CI: 1.54–7.15; p-value<0.001) and in those living in balanced houses (OR: 2.64; 95% CI: 1.57–4.43; p-value<0.001) compared with people living in under-occupied houses (Table 4). The proportion of variation at the household level (ICC) was 74.0%.

Sensitivity analyses were consistent with the finding that overcrowding was associated with increased risk of PCR and anti-N total immunoglobulin assay antibody confirmed SARS-CoV-2 (Table 5). In a sensitivity analysis where we did not subtract 1

| Characteristic               | Univariable logistic regression | Multivariable logistic regression |
|------------------------------|---------------------------------|----------------------------------|
|                              | OR     | 95% CI    | p-value | OR     | 95% CI    | p-value |
| Ethnicity                    |        |           |         |        |           |         |
| White British                | 1.00   | —         | —       | 1.00   | —         | —       |
| White Irish                  | 0.82   | (0.39, 1.72) | 0.60   | 0.78   | (0.28, 2.14) | 0.63   |
| White Other                  | 0.86   | (0.56, 1.31) | 0.48   | 0.44   | (0.24, 0.80) | 0.007  |
| Mixed                        | 1.51   | (0.85, 2.69) | 0.16   | 0.86   | (0.44, 1.68) | 0.66   |
| South Asian                  | 2.05   | (1.16, 3.62) | 0.013  | 1.00   | (0.49, 2.04) | 0.99   |
| Other Asian                  | 0.89   | (0.33, 2.44) | 0.83   | 0.50   | (0.14, 1.80) | 0.29   |
| Black                        | 1.25   | (0.37, 4.30) | 0.72   | 0.66   | (0.14, 3.10) | 0.60   |
| Other Ethnicity              | 1.12   | (0.32, 3.87) | 0.859  | 0.58   | (0.12, 2.81) | 0.50   |
| Missing                      | 0.71   | (0.31, 1.62) | 0.418  | 0.13   | (0.00, 6.53) | 0.305  |
| Household income combined for last 12 months | | | | | | |
| £0-9,999                     | 1.28   | (0.76, 2.16) | 0.36   | 0.98   | (0.47, 2.06) | 0.964  |
| £10,000-24,999               | 1.00   | —         | —       | 1.00   | —         | —       |
| £25,000-49,999               | 0.94   | (0.71, 1.25) | 0.668  | 0.81   | (0.54, 1.23) | 0.332  |
| £50,000-£74,999              | 1.38   | (1.00, 1.91) | 0.05   | 0.91   | (0.58, 1.43) | 0.68   |
| £75,000-£99,999              | 1.26   | (0.84, 1.90) | 0.27   | 0.73   | (0.42, 1.29) | 0.28   |
| £100,000+                    | 1.60   | (1.07, 2.40) | 0.02   | 0.90   | (0.52, 1.54) | 0.693  |
| Prefer not to say            | 0.74   | (0.45, 1.23) | 0.249  | 0.59   | (0.28, 1.25) | 0.168  |
| Missing                      | 1.20   | (0.84, 1.72) | 0.315  | 0.99   | (0.59, 1.64) | 0.957  |
| Geographical region          |        |           |         |        |           |         |
| East of England              | 1.00   | —         | —       | 1.00   | —         | —       |
| East Midlands                | 1.31   | (0.90, 1.91) | 0.15   | 1.32   | (0.78, 2.22) | 0.30   |
| London                       | 1.89   | (1.37, 2.62) | <0.001 | 1.48   | (0.93, 2.37) | 0.098  |
| North East                   | 1.49   | (0.94, 2.37) | 0.09   | 1.51   | (0.81, 2.83) | 0.20   |
| North West                   | 1.66   | (1.17, 2.34) | 0.004  | 1.68   | (1.05, 2.69) | 0.032  |
| South East                   | 0.97   | (0.71, 1.32) | 0.845  | 0.95   | (0.61, 1.48) | 0.820  |
| South West                   | 0.79   | (0.51, 1.21) | 0.27   | 0.81   | (0.43, 1.55) | 0.53   |
| West Midlands                | 1.05   | (0.66, 1.66) | 0.85   | 1.08   | (0.55, 2.09) | 0.83   |
| Yorkshire and The Humber     | 1.91   | (1.22, 2.98) | 0.005  | 2.04   | (1.13, 3.70) | 0.019  |
Table 4. Univariable and multivariable mixed effects logistic regression models examining the association between household overcrowding and SARS-CoV-2 antibody test results.

| Characteristic                        | Univariable logistic regression | Multivariable logistic regression |
|---------------------------------------|---------------------------------|----------------------------------|
|                                       | OR     | 95% CI | p-value | OR     | 95% CI | p-value |
| Household overcrowding category       |        |        |         |        |        |         |
| Under occupied                        | 1.00   | —      | —       | 1.00   | —      | —       |
| Balanced                              | 5.01   | (3.13, 8.01) | <0.001 | 2.64   | (1.57, 4.43) | <0.001 |
| Overcrowded                           | 8.20   | (4.04, 16.66) | <0.001 | 3.32   | (1.54, 7.15) | <0.001 |
| Age group                             |        |        |         |        |        |         |
| 18–24                                 | 9.25   | (5.02, 17.08) | <0.001 | 7.39   | (3.97, 13.77) | <0.001 |
| 25–44                                 | 6.75   | (4.58, 9.95) | <0.001 | 3.92   | (2.58, 5.94) | <0.001 |
| 45–64                                 | 3.40   | (2.54, 4.54) | <0.001 | 3.00   | (2.24, 4.02) | <0.001 |
| 65+                                   | 1.00   | —      | —       | 1.00   | —      | —       |
| Sex                                   |        |        |         |        |        |         |
| Male                                  | 1.00   | —      | —       | 1.00   | —      | —       |
| Female                                | 0.94   | (0.78, 1.13) | 0.517  | 0.87   | (0.72, 1.05) | 0.15  |
| Ethnicity                             |        |        |         |        |        |         |
| White British                         | 1.00   | —      | —       | 1.00   | —      | —       |
| White Irish                           | 1.18   | (0.49, 2.86) | 0.708  | 0.82   | (0.34, 2.02) | 0.67  |
| White Other                           | 2.32   | (1.41, 3.82) | 0.001  | 0.71   | (0.40, 1.24) | 0.23  |
| Mixed                                 | 3.23   | (1.25, 8.37) | 0.016  | 1.40   | (0.52, 3.81) | 0.51  |
| South Asian                           | 5.38   | (2.19, 13.19) | 0.000  | 1.53   | (0.58, 4.0)  | 0.39  |
| Other Asian                           | 0.77   | (0.15, 3.93) | 0.750  | 0.29   | (0.05, 1.50) | 0.14  |
| Black                                 | 3.35   | (0.54, 20.72) | 0.193  | 0.73   | (0.10, 5.18) | 0.76  |
| Other Ethnicity                       | 0.70   | (0.08, 5.90) | 0.745  | 0.26   | (0.03, 2.32) | 0.23  |
| Household income combined for last 12 months |        |        |         |        |        |         |
| £0–9,999                              | 1.99   | (0.97, 4.07) | 0.060  | 1.40   | (0.69, 2.85) | 0.35  |
| £10,000–24,999                        | 1.00   | —      | —       | 1.00   | —      | —       |
| £25,000–49,999                        | 1.32   | (0.90, 1.95) | 0.156  | 1.10   | (0.75, 1.61) | 0.62  |
| £50,000–£74,999                       | 2.62   | (1.68, 4.08) | 0.000  | 1.56   | (1.00, 2.44) | 0.05  |
| £75,000–£99,999                       | 2.70   | (1.54, 4.74) | 0.001  | 1.31   | (0.74, 2.33) | 0.36  |
| £100,000+                             | 2.87   | (1.63, 5.03) | 0.000  | 1.21   | (0.68, 2.18) | 0.52  |
| Prefer not to say                     | 0.92   | (0.47, 1.82) | 0.812  | 0.64   | (0.33, 1.25) | 0.19  |
| Missing                               | 1.96   | (1.11, 3.44) | 0.020  | 1.28   | (0.73, 2.26) | 0.39  |
| Geographical region                   |        |        |         |        |        |         |
| East of England                       | 1.00   | —      | —       | 1.00   | —      | —       |
| East Midlands                         | 1.28   | (0.75, 2.18) | 0.366  | 1.27   | (0.74, 2.17) | 0.39  |
from the number of rooms there was evidence of an increased odds of having a positive SARS-CoV-2 antigen result in balanced (OR: 2.08; 95% CI: 1.16–3.73; p-value=0.015) and overcrowded households (OR: 2.37; 95% CI: 1.16–4.83; p-value=0.017) compared with those in under-occupied houses.

### Discussion

We estimate that overcrowded households have between 2- and 4-times the odds of PCR-confirmed SARS-CoV-2 and anti-N total immunoglobulin assay antibody positivity compared to under-occupied households. These elevated odds of confirmed
SARS-CoV-2 remained after we adjusted for differences in demographic and other socio-economic factors between households that may increase risk of being infected. We also found that people in accommodation considered balanced, where the number of rooms was equal to the number of people, were at increased risk of PCR confirmed SARS-CoV-2 compared to under-occupied houses.

Virus Watch is a large national household community cohort study of the occurrence and risk factors for SARS-CoV-2 infection and is designed to estimate incidence of PCR-confirmed SARS-CoV-2 and measure effectiveness and impact of recommended COVID-19 control measures. Individuals in the study are geographically distributed across England and Wales and the cohort is diverse in terms of age, sex, ethnicity and socio-economic composition, and levels of overcrowding were comparable with national estimates for England and Wales. Respondents to our housing survey were more likely to be White British, over the age of 65 and have a higher income. As a result, our unadjusted estimates of the association between overcrowding and SARS-CoV-2 are likely to underestimate the strength of the association, as minority ethnic households have the highest rates of overcrowding. Our multivariable estimates adjust for age, ethnicity and income; however, due to the upper limit of 6 people in a household and the higher socio-economic status of participants in Virus Watch compared with the general population, our estimates are likely to underestimate the true magnitude of the association.

Several additional limitations are important when interpreting our findings. Virus Watch participants were required to reside at least four days at the same address, but as Virus Watch did not collect more detailed information on time spent with other household members inside, more granular data on intra-household contacts could be collected in the future to examine this issue further. Individuals living in nursing or residential homes were not eligible for inclusion in Virus Watch and, therefore, our results are not generalisable to these settings. The linked PCR data we included from the Second Generation Surveillance System for SARS-CoV-2 did not include lateral-flow device results and as such we will miss individuals who tested at home without a confirmatory PCR test, but note that anyone testing positive with a lateral-flow test is advised to have a confirmatory PCR test.

There were more participants taking part in the February 2021 monthly survey that were over the age of 65 compared with the general population in England and Wales, more people of White British ethnicity, and more people of household size of 2. Virus Watch is limited by the fact that only households with a lead household member able to speak English and access the internet were able to take part in the study. An important additional limitation is that only households of up to six people were eligible for inclusion. By not including households of over six people, we are likely underestimating the risk associated with overcrowding. Conversely, we cannot rule out residual confounding as a possible alternative explanation for some of the excess risk in the associations we describe.

Access to SARS-CoV-2 antigen testing is socio-economically patterned, with those in more deprived areas having less ability to access tests and less likely to be contact traced. Our data may therefore under-estimate the true infection risk associated with overcrowded housing. To mitigate this bias in access to antigen testing, we analysed antibody data from the laboratory cohort where home test kits were provided to all adult participants.

Our findings highlight the importance of public health interventions to prevent and stop the spread of SARS-CoV-2 considering the much greater risk of being infected for people living in overcrowded households. The pathways between overcrowding and increased risk of infection are not fully understood. Overcrowded households are more likely to be found in socio-economically deprived and urban areas, and higher levels of COVID-19 in these communities include higher likelihood of working in public-facing and essential occupations, lower likelihood of working from home during lockdowns, constraints to self-isolation when ill or in contact with a case (both financial and due to lack of suitable space within an overcrowded house), increased number of household members sharing common spaces and facilities. Overcrowded households are also more likely to suffer from environmental exposures that have been linked to COVID-19 including air pollution.

Overcrowding is an important risk through which COVID-19 inequalities manifest. To address social, ethnic and regional inequalities in COVID-19 outcomes, public health responses should explicitly address overcrowded housing. Infectious cases in overcrowded houses are less likely to be able to be isolated from other household members or to avoid using shared spaces, including bathrooms. However, existing advice on the importance of infectious cases using masks and ventilation through opening of windows should be supported by tailored communications strategies. Schemes that provide hotel accommodation for cases with vulnerable residents in overcrowded households, such as those used internationally and recently introduced in the London Borough of Newham, should be considered a priority public-health intervention. Such community quarantine and isolation options require clear, inclusive guidance with specific advice for large households and multigenerational families. Although the UK does not target COVID-19 vaccine according to social factors, our results highlight the need to ensure high vaccine coverage in areas with high levels of overcrowding. There is currently evidence of lower COVID-19 vaccination rates in such areas, highlighting the importance of working closely with local community groups to increase vaccination intention and ensuring vaccination is highly accessible.

Improving the standard and supply of housing in the UK is important for the COVID-19 recovery as 32% of all households are estimated to experience overcrowding, affordability challenges or non-decent housing that influence their risk. In relation to these challenges, a paper by the ethnicity sub-group of the Scientific Advisory Group for Emergencies (SAGE) recommends the following measures to reduce the risk of transmission...
within households: provision of emergency grants for repair and maintenance of social and private rental housing, particularly to increase ventilation; removing the benefit cap and welfare restrictions, particularly in high housing cost areas and for immigrant families; reviewing implementation of the bedroom tax for those in multi-person households; and investment in affordable childcare and alternative community spaces for social connection, particularly for the elderly.

Addressing England’s overcrowding challenge is complex and requires action from multiple government and private sector stakeholders, including the following: revising the statutory overcrowding standard, increasing the housing supply and size of homes/rooms, and reducing under-occupation in social housing. Breaching the statutory overcrowding standard is a criminal offence, yet the threshold is considered to be too low such that very few homes are statutorily overcrowded, reducing local authorities’ ability to take action. Housing supply has fallen short of demand for decades, which is caused by inter-related factors, such as construction capacity, planning regulation, public opposition to new development, overseas buyers, overreliance on the private sector and underuse of social approaches to housing development. English homes are among the smallest in Europe and the Nationally Described Space Standard for new homes is only mandatory if implemented through local planning policies following tests for need and viability. Finally, although there are under-occupied and empty homes in England, these are not suitable to solve the housing supply problem due to their location, ownership and other factors.

Housing is an important determinant of health for many physical and mental health conditions and COVID-19 has exposed, in real-time, vulnerabilities in the English housing stock, particularly for low-income and minority ethnicity populations. Looking forward, the country’s health and resilience to future pandemics and other risks, such as climate change, are dependent on the state of the housing stock. This analysis of overcrowding and COVID-19 underscores the need to improve housing for health. As part of the agenda to “Build Back Better” and fairer, investment in sustainable, high-quality and more affordable housing will support health, jobs and the wider economic recovery so that we are prepared for the key challenges of the 21st century.

Data availability
Underlying data
We aim to share aggregate data from this project on our website and via a “Findings so far” section on our website - https://ucl-virus-watch.net/. We are sharing individual record level data on the Office of National Statistics Secure Research Service, and given the sensitive content in our dataset (information on income and household characteristic) for this study, we currently cannot release the data at the individual level. In sharing the data we will work within the principles set out in the UKRI Guidance on best practice in the management of research data. Access to use of the data whilst research is being conducted will be managed by the Chief Investigators (ACH and RWA) in accordance with the principles set out in the UKRI guidance on best practice in the management of research data. Data access requests to data can also be made directly to the Virus Watch chief investigators (ACH or RWA) at the following email address: viruswatch@ucl.ac.uk.

Extended data
Zenodo: Extended data for ‘Household overcrowding and risk of SARS-CoV-2: analysis of the Virus Watch prospective community cohort study in England and Wales’ https://doi.org/10.5281/zenodo.5708766

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Author contributions
Conceptualization (RWA, HP, EF, JK, VN, SB, TB, AA, CS, DD, JT, SVK, WLEF, CG, PP, MS, IB, AMDN, AH), Data Curation (RWA, EF, VN, SB, TB, WLEF, CG, AMDN), Formal Analysis (RWA,ME), Investigation (All), Methodology (All), Project Administration (RWA, JK), Resources (All), Writing – Original Draft Preparation (RWA and HP), Writing – Review & Editing (All).

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Eben Kenah
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It is great to see the implementation of a new generation of Virus Watch studies. Although I have several comments about the statistical analysis, the overall conclusions of the study appear sound. Its most important limitations are discussed adequately, and its implications for public health are clearly explained.

- Overcrowding measure: The categorization of the number of people per room is not sufficiently justified. Why not use the original variable and fit a flexible model using orthogonal polynomials or splines? Odds ratios could be reported comparing specific levels of the people/room measure, which might be easier to interpret than the comparisons of overcrowded and balanced households with und-occupied housing. The under-occupied and overcrowded categories are mixtures of households at different levels of crowding.

- Covariates: The number of individuals in the household is likely to act as a confounder of the association between overcrowding and the risk of SARS-CoV-2 infection. Even if we imagine that the transmission probability between household members is constant in any given crowding category, a larger household can support longer chains of transmission and a higher risk of infection from within-household transmission. Although some of this confounding may be attenuated by the household-level random effect, it would be better to include household size as a covariate. The resulting model would isolate crowding as an exposure more effectively. However, I doubt that this would affect the qualitative conclusion that crowding is associated with higher risk of infection.

- Results: Given that the vast majority of the study population lives in "under-occupied" housing, it seems like the categories used in the study might not be a great fit for the context. Besides using the original people/rooms variable (see above), another alternative would be to base the categories on quantiles of the crowding measure. I acknowledge that the categories reported here match those used in earlier research on overcrowding, but I am not sure they are the best choice for the question being addressed here.
- Figure S1: The explicit use of a DAG to identify possible confounders is commendable, but the DAG would be much easier to read and interpret if a consistent direction of time was maintained (e.g., left to right or top to bottom). The DAG has arrows pointing left, right, up, and down.

- R code: The code is well-formatted and readable. My only comment is that profile or bootstrap confidence intervals are generally better choices than Wald confidence intervals. Is there a reason that these were not used?

Is the work clearly and accurately presented and does it cite the current literature?
Yes

Is the study design appropriate and is the work technically sound?
Yes

Are sufficient details of methods and analysis provided to allow replication by others?
Yes

If applicable, is the statistical analysis and its interpretation appropriate?
Partly

Are all the source data underlying the results available to ensure full reproducibility?
Yes

Are the conclusions drawn adequately supported by the results?
Yes

Competing Interests: No competing interests were disclosed.

Reviewer Expertise: epidemiologic methods, survival analysis, causal inference, statistical methods for infectious disease epidemiology, epidemic models, networks

I confirm that I have read this submission and believe that I have an appropriate level of expertise to confirm that it is of an acceptable scientific standard.

Kathy Leung
The University of Hong Kong, Hong Kong, Hong Kong

The authors presented an analysis of SARS-CoV-2 infection risk using data from the Virus Watch
study in the UK. The design of the Virus Watch study is robust, and it is an advantage that the
survey results can be linked to the SARS-CoV-2 testing results in the national database. However, I
have reservations about the design and selection of regression model of this study. Please see
below for my major comments:

1. The definition of overcrowding relied only on the calculation of the number of household
occupants divided by the number of rooms but did not consider the composition of
household occupants. For example, households with school-aged children tend to be more
crowded, and school-aged children are considered to be the driver of transmission in many
respiratory infectious diseases. It is therefore difficult to separate the effects of
overcrowding and having one or more school-aged children in the households.
2. Similarly, the study was carried out between June 2020 and February 2021 and the UK
experienced several extended periods of lock downs. I presume lock down would increase
the time residents spent at home but there should be a differential effect for people with
different working statuses. The occupations of household members and whether they were
able to work from home during lock downs should be considered in the model.
3. Since the mass vaccination programme started in late 2020 and/or early 2021, the model
should also consider the differential vaccination coverage across different age groups. The
older adults aged 65 or above were prioritised in early 2021 in the vaccination programme.

**Is the work clearly and accurately presented and does it cite the current literature?**
Yes

**Is the study design appropriate and is the work technically sound?**
Yes

**Are sufficient details of methods and analysis provided to allow replication by others?**
Yes

**If applicable, is the statistical analysis and its interpretation appropriate?**
Partly

**Are all the source data underlying the results available to ensure full reproducibility?**
Yes

**Are the conclusions drawn adequately supported by the results?**
Partly

**Competing Interests:** No competing interests were disclosed.

**Reviewer Expertise:** infectious disease modelling

I confirm that I have read this submission and believe that I have an appropriate level of
expertise to confirm that it is of an acceptable scientific standard, however I have
significant reservations, as outlined above.
Melissa Rolfes
Influenza Division, Centers for Disease Control and Prevention, Atlanta, GA, USA

Summary:
The authors describe an analysis of SARS-CoV-2 infection risk among a community cohort in the UK. This is a long-standing cohort with study procedures and activities that pre-date the COVID-19 pandemic. Participants who completed a monthly survey in February 2021 were considered for analysis. These participants provided demographic and household characteristics as well as vaccination information. Household information was used to classify participants as living in under-occupied, balanced, or overcrowded settings based on the number of people per rooms in the household. The risk of SARS-CoV-2 infection was compared between these different living situations using two outcomes: 1) evidence of acute SARS-CoV-2 infection using information from linked national COVID-19 testing information systems and 2) evidence of acute infection from anti-N antibodies collected from a sub-set of participants. Using these two outcomes the authors found that participants living in over-crowded settings (with greater numbers of people per room) and those living in balanced settings (with equal numbers of people per room) had significantly greater odds of SARS-CoV-2 infection. This evidence stood up to several sensitivity analyses which examined the impact of assumptions and potential confounders.

The authors conclude that there should be greater investment in reducing the occurrence of overcrowding and also in providing means to prevent and control transmission of infectious diseases within overcrowded settings. This type of investment may need to be large but also innovative as individuals living in overcrowded settings are at additional social and economic disadvantages that impact uptake and success of common public health measures.

Major comments:
- Background would benefit from a statement or two about what this study provides that prior literature doesn’t. The authors point to many studies (retrospective and prospective) to demonstrate greater incidence of COVID-19 in households with a greater density of people. I think it would be helpful to highlight what gaps this study and analysis fill.

- In the first paragraph of the Methods section, it would be helpful to state the study procedures at a high-level. There are very detailed study procedures in other publications, but a few broad points would be helpful here, too. I was looking for information about inclusion/exclusion criteria, duration of recruitment (recruitment began in June 2020, when did it end or is it rolling?), anticipated duration of follow-up when participants enroll, etc. The main webpage for the Virus Watch study has a nice, concise summary that has these details. I strongly encourage the authors to modify this Methods paragraph to include additional information.
  - The Discussion section does include some inclusion/exclusion criteria, but I strongly encourage the authors to move those details to the Methods.

- Additionally, in the methods, it would be helpful to know if participants needed to enroll
along with other household members or if participants could enroll alone (even if their other household members did not enroll). I think adding in the inclusion/exclusion criteria would be helpful to address this confusion.

- This information also has relevance for the interpretation of the household characteristics in Table 1 and Table 2. If multiple people per household were allowed to participate, then there is a correlation induced in the presented proportions and would not necessarily be able to be compared with the ONS summary data.

- Some readers may not be familiar with the UK test and trace program. Please include in the Methods (rather than in the Discussion) a few details about the type of laboratory assay that was be conducted. was it a molecular/PCR assay or a rapid antigen assay? If it was a rapid antigen test, were positives confirmed by PCR? There is some use of “PCR” and some use of “antigen test” in the Results section that could create confusion without some further details.

- It is not clear whether households with just 1 member were included in analysis. It seems like including participants who do not have another household member sharing their living space would have different risk from different places than those who did share their living space. Suggest including information in the Methods about how single-person households were managed in the analysis. And if they were original included, I suggest the authors include a sensitivity analysis excluding the single-person households.

- The authors describe a sub-cohort of ~10,000 participants who provided capillary blood specimens for antibody testing. It would be helpful for readers to see a column in Table 1 describing these 10,000 participants in the sub-cohort. This level of detail would help the reader assess confounding, as the magnitude of the association between SARS-CoV-2 and overcrowding was greater when the outcome was antibody results. As a reader, I wonder whether the demographic and household characteristics differed in the sub-cohort and could account for some of the differences in the magnitude of the association with overcrowding.

- In Table 2, the results of the PCR tests are reported, by household and demographic characteristics. The PCR-negative and not tested group are combined into one, however, this “lumping” is potentially misleading. There are often social and demographic features associated with getting tested for SARS-CoV-2 and those associations would be masked if these two groups are lumped together. I suggest the authors consider making 3 columns for the PCR test results in Table 2: Not tested | Negative | Positive. If it is not possible to separate these two groups (negative and not tested) because of the way that the test and trace program reporting was done, please add that detail to the Methods to inform readers of the reason behind lumping.

- Furthermore, for the regression analysis, the decision could be made to lump negative with not tested, but the reader isn't able to assess the viability of that lumping without seeing the two groups split in Table 1.

- The authors describe that information on COVID-19 vaccination was collected from participants but have not included vaccination in the analysis. Given that vaccines during this time period (Feb–Aug 2021) were well-matched to the circulating SARS-CoV-2 variant and provided meaningful protection against infection and that vaccination uptake could be
associated with overcrowding (as the authors point out in the Discuss), I strongly encourage the authors to include vaccination as a covariate in the regression models.

Minor comments
○ In Table 1, ONS is used as an acronym but doesn’t seem to have been previously defined.

○ In Tables 3 and 4, the p-values are presented with inconsistent significant digits but would be helpful to be consistent.

○ In the Discussion, there are several links to websites with information on public health control measures and further guidance. These should be included as numbered references rather than as hyperlinks within the text. This change would enable those readers with only a text pdf or limited internet access to reach these linked urls.

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Yes

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Partly

Are sufficient details of methods and analysis provided to allow replication by others?
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If applicable, is the statistical analysis and its interpretation appropriate?
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Are all the source data underlying the results available to ensure full reproducibility?
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Are the conclusions drawn adequately supported by the results?
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Competing Interests: No competing interests were disclosed.

Reviewer Expertise: Epidemiology, influenza, transmission

I confirm that I have read this submission and believe that I have an appropriate level of expertise to confirm that it is of an acceptable scientific standard, however I have significant reservations, as outlined above.