Feeding on Multiple Sources: Towards a Universal Parameterization of the Functional Response of a Generalist Predator Allowing for Switching

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Abstract

Understanding of complex trophic interactions in ecosystems requires correct descriptions of the rate at which predators consume a variety of different prey species. Field and laboratory data on multispecies communities are rarely sufficient and usually cannot provide an unambiguous test for the theory. As a result, the conventional way of constructing a multi-prey functional response is speculative, and often based on assumptions that are difficult to verify. Predator responses allowing for prey selectivity and active switching are thought to be more biologically relevant compared to the standard proportion-based consumption. However, here we argue that the functional responses with switching may not be applicable to communities with a broad spectrum of resource types. We formulate a set of general rules that a biologically sound parameterization of a predator functional response should satisfy, and show that all existing formulations for the multispecies response with prey selectivity and switching fail to do so. Finally, we propose a universal framework for parameterization of a multi-prey functional response by combining patterns of food selectivity and proportion-based feeding.

Introduction

Population dynamics of multi-species communities is a major challenge in contemporary ecology [1–5]. In particular, parameterization of the functional response of a generalist predator feeding on multiple resources has been a focus of ecological literature for a few decades already and various approaches have been suggested [6–12]. It has previously been shown that different forms of the response can result in different predictions for the dynamics of the community [8], [13], [14]; therefore, an understanding of this issue is central to community and food web ecology. However, while the dynamics of few-species systems have by now been studied almost comprehensively [but see [15], [16]], good understanding of multi-species systems (that can include hundreds or even thousands of species) is still lacking.

Parameterization of a predator functional response for an arbitrary number of food sources can be derived using various theoretical approaches. For instance, one can utilize the optimal foraging theory framework [9], [10], [17–22]; it can be based on specific biological traits such as the predator’s memory [23] or simply use common-sense reasoning [11], [12], [14], [24–26]. Empirical verification of different parameterizations, however, remains a considerable challenge: although experimental data on feeding of predator/consumers on multiple resources are abundant both in marine and terrestrial ecology [27–34], laboratory experiments and field observations are mostly limited to the case where the predator has the choice of only a few varieties of food (n ≤ 4). On the other hand, ecosystems often contain a large number of species belonging to the same trophic level. Moreover, even within a single prey population individuals can largely differ in terms of size/age and behavior [e.g. [35]] so that, from the predator’s point of view, they are likely to be different food sources. There can be also a pronounced variation in terms of the parasite load that organisms carry and predation can heavily depend on the degree of infectivity of prey [36]. Thus, in real ecosystems there can be hundreds or thousands of food sources which are, from the predator’s perspective, either similar or distinctly different.

Construction of a multi-prey functional response in the presence of a large variety of resources is not straightforward, as it always requires some hypotheses about the predator foraging behavior that can be species-specific and are very difficult or even impossible to test experimentally (cf. [11]). This uncertainty has resulted in a broad variety of different mathematical formulations of functional response, each them having some rationale behind the equations; see [8] and [11] for a review. The type of response which is regarded as simplest is the so-called proportion-based consumption where individual predators pick up the prey items randomly and thus the food intake rate is determined by the relative frequency of the prey species in the community [11], [37], [38]. However, the proportion-based response is often thought to miss the complexity of foraging behavior [6], [8]. Eventually, more sophisticated parameterizations were developed to describe food intake by animals with complex foraging strategies, e.g. exhibiting food selectivity and active switching. In this case, a predator can ‘switch’ towards a more energetically profitable food source,
The paper is structured as follows. In Section 2, we introduce a set of rules that the multi-prey functional response of a predator should satisfy. In Section 3, we carefully examine the existing formulations of a multi-prey functional response and show that none of them satisfies these rules. In Section 4, we show how the drawbacks of the previous formulations of functional responses can be amended, resulting in a transitional parameterization of the predator’s intake rate which describes consumption of prey with both close and distinctly different life history traits. In Section 4, we discuss the implications of our findings; in particular, we show how are results may contribute to the recent efforts in the literature to relate the biodiversity and productivity of ecosystems.

Results

Basic rules for a multi-prey functional response

Throughout this paper, we assume that the predator functional response is prey-dependent. A known alternative is the ratio-dependent response that takes into account interference between predators [53], [54]. However, inter-predator interference is likely to become important only when the predator population density is very high. Considering the plankton community as a paradigm of a multispecies system, phytoplankton (i.e. prey) outbreaks are well known and relatively frequent phenomena, cf. “red tides” and “green tides” as well as seasonal phytoplankton blooming [47]. On the contrary, mesozooplankton (predator) outbreaks up to the densities where the presence of other foragers could strongly affect individual intake rate are very rare (e.g. [55]). We therefore restrict our analysis to the cases when the predator density is not very high without any substantial loss of generality.

We consider that the single resource functional response (i.e. in the absence of all other resources) is described by the function \( f_0(P_i) \), where \( P_i \) is the density of resource \( i \). Function \( f_0(P_i) \) therefore describes the per capita rate of increase of a given ‘consumer’ species when it feeds on the resource \( i \) only. This rate of increase can be affected by the presence of other resources. Correspondingly, we denote the multiple resource functional response (describing the intake rate of resource \( i \) in the presence of other resources) by \( f_i(\overline{P}) \), where \( \overline{P} = (P_1, P_2, \ldots, P_n) \).

In the below, we will refer to the consumer species as predator and to the resource species as prey. We shall introduce several criteria or rules that are meant to account for biologically reasonable properties of the predator functional response and to avoid artificial or unrealistic behavior. In order to reveal any potential artifacts, we apply a special test: we consider a thought experiment where the given prey species is split into several groups or subpopulations in such a way that an observer can distinguish between the groups but the predator cannot. In reality, such splitting could be done by using certain chemical markers or radioactive tracers. The idea of this thought experiment is that, since all those groups consist of essentially the same prey species, a ‘good’ functional response should be stable with respect to such splitting while a ‘bad’ one would not.

We begin with the rule that is known in the ecological literature as the consistency requirement condition [23], [56]. Here we use it in the following form:

**Rule (i).** A biologically reasonable parameterization for the multi-prey functional response \( f_i(\overline{P}) \) should be applicable to the case where all the resources have similar properties.

Here “similar properties” means that, ultimately, different resources may be indistinguishable from each other. Following the idea of our thought experiment, let us consider that a single prey species is divided arbitrarily into \( N \) groups. Obviously, each of
these groups has exactly the same parameters and hence their consumption by the predator ought to be described by the same function. In this case, the following identity should apply:

\[ f_0(P) = \sum_i f_i(\bar{P}), \quad P = \sum_i P_i. \]  
(1)

In other words, for a sound functional response splitting the same species into several groups and summing back the consumption rates should give the same consumption rate as for the single species with the overall density \( P \). Note that it does not matter whether the prey species in different groups is actually the same or only “similar” as long as the predator cannot spot the difference. Therefore, Rule (i) should apply as well in the situation, where the prey species are morphologically different but from the predator perception they are similar as potential food resources.

Rule (i) addresses the case of feeding on a single prey species or on several prey species with similar traits. Generally speaking, the presence of alternative resources with distinctly different traits can change the situation. However, similarly to the above, the predator’s food consumption of a given resource (prey) should remain stable with respect to its division. Hence we arrive at the following rule:

**Rule (ii).** Parameterization \( f_i(\bar{P}) \) should remain valid when resource \( i \) is split into \( m \) groups in the presence of other resources.

Rule (ii) therefore gives an extension of Rule (i). As we will show below, this extension is nontrivial and goes beyond the basic consistency requirement condition. Splitting the \( i \)th resource into \( m \) groups can reflect some actual intraspecific variation of certain life traits but it can just as well be a part of the thought experiment. In order to obtain a mathematical formulation of Rule (ii), we denote the resource partition in the system before splitting as \( \bar{P} \), and after splitting as \( \bar{P}_1, \ldots, \bar{P}_m \), with \( P_{ij} \) being the density of each group of resource \( i \). The total intake rate of resource \( i \) (now consisting of \( m \) groups) should then satisfy the following equation:

\[ f_i(\bar{P}) = \sum_{j=1}^m f_{ij}(\bar{P}_j), \]  
(2)

or, equivalently,

\[ f_i(P_1, P_2, \ldots, P_m, \ldots, P_n) = \sum_{j=1}^m f_{ij}(P_1, P_2, \ldots, P_i, P_{i2}, \ldots, P_{im}, \ldots, P_n). \]

Therefore, splitting the \( i \)th prey species into \( m \) groups – now in the presence of alternative prey species – and summing back the consumption rates over the groups should give the original consumption rate for the resource \( P \). We want to emphasize that satisfaction of the logical consistency requirement as given by Rule (i) does not, in general, guarantee satisfaction of Rule (ii). In Section 3 we will show an example of such situation.

**Rule (iii).** Preyation on any two species with close life traits cannot change the ratio of their population densities, which will therefore remain constant over time.

Indeed, consider the situation when a predatory species feeds on two prey species with close life traits. For the sake of simplicity we assume that there is only one predator. (Following the same approach as above, ultimately, we can consider the population of the same prey species with population density \( P \) split into two groups with densities \( P_1 \) and \( P_2 \) respectively.) In this case, the predator cannot distinguish between the groups and hence it is bound to proportionate consumption. Correspondingly, the initial ratio of \( P_1 \) and \( P_2 \) should remain the same over time, i.e. \( P_1/P_2 \approx \text{const} \), which we can write as

\[ \frac{d}{dt} \left( \frac{P_1}{P_2} \right) = \frac{P_2 \dot{P}_1 - P_1 \dot{P}_2}{P_2^2} \approx 0 \iff P_2 \dot{P}_1 - P_1 \dot{P}_2 \approx 0, \]  
(3)

The temporal dynamics of each of these species can be described by the following equation which is a standard equation in food web models (e.g. [57] also see Eqs.6–8 below):

\[ \frac{dP_i}{dt} = P_i \left( r_i - \mu_i - f_i(\bar{P}) Z \right), \]  
(4)

where \( r_i \) and \( \mu_i \) \( (i = 1, 2) \) are the growth and mortality rates of prey and \( Z \) is the density of predator. Equation (4) accounts for the fact that the change in the prey density is due to growth, mortality and predation. Since the two prey species are assumed to have similar traits, their per capita growth rates and mortalities have close values, i.e. \( r_1 \approx r_2 \) and \( \mu_1 \approx \mu_2 \). Using (4), we then can re-write condition (3) as

\[ P_2 f_1(\bar{P}) - P_1 f_2(\bar{P}) \approx 0. \]  
(5)

Equation (5) thus gives a mathematical expression of Rule (iii).

In a more general case of \( N \) species with close traits (or by splitting the whole population into \( N \) parts), we can easily derive a condition analogous to (5) for any pair of \( P_i \) and \( P_j \), \( 1 \leq i, j \leq N \). Expression (5) should therefore hold for all prey species that are different but have close life traits; in particular, when \( r_i \approx r_j \) and \( \mu_i \approx \mu_j \).

**Rule (iv).** The total predicted biomass of prey and/or predator in generic food chain models should not be sensitive to the way in which the species are divided into groups.

Let us consider a generic food web consisting of a large number of competing prey species. A standard model is given by the following equations (cf. [12], [57], [58]):

\[ \frac{dN_i}{dt} = D_0 (N_0 - N_i) - \sum_j r_i \left( \bar{N} \right) P_j, \]  
(6)

\[ \frac{dP_i}{dt} = P_i \left( r_j \left( \bar{N} \right) - \mu_j \right) - f_i(\bar{P}) Z, \]  
(7)

\[ \frac{dZ}{dt} = Z \left( \theta \sum_i f_i(\bar{P}) - \delta \right), \]  
(8)

where \( N_i \) is the density of the limiting nutrients; \( P_i \) are the densities of prey (e.g. phytoplankton), \( i = 1, \ldots, n \quad (n \gg 1) \); \( Z \) is the density of predator (e.g. zooplankton). For the sake of simplicity we consider that there is only one predator. Parameter \( \theta \) is the food utilization coefficient, \( \delta \) is the mortality of the predator, \( r_i \) and \( \mu_i \) are the growth rate and mortality of the prey species \( i \), respectively, and \( D_0 \).
Table 1. Results of testing the existing multi-prey predator’s functional responses against the set of the basic Rules (i)–(vi). For details, see the text and the online appendix in Material S1.

| Functional response, reference | Parameterization | Rule (i) | Rule (ii) | Rule (iii) | Rule (iv) | Rule (v) | Rule (vi) (Holling type for a single resource) |
|-------------------------------|-----------------|----------|-----------|------------|-----------|---------|---------------------------------------------|
| [12], [14] and [64]          | $g_{\text{max}} \frac{\rho_i P_j}{K + \sum_i \rho_i P_j}$ | NO       | NO        | NO         | NO        | NO      | Type II                                     |
|                              |                 |          |           |            |           |         |                                             |
| [26]                         | $g_{\text{max}} \frac{\rho_i P_j}{K + \sum_i \rho_i P_j}$ | NO       | NO        | NO         | NO        | YES     | Type III (sigmoid)                          |
|                              |                 |          |           |            |           |         |                                             |
| [24]                         | $g \frac{\rho_i P_j}{\sum_i \rho_i P_j}$ | NO       | NO        | NO         | NO        | NO      | Type I (can be extended to type II)         |
|                              | $P_{ij}>1$           |          |           |            |           |         |                                             |
| [62]                         | $g \frac{\rho_i P_j}{\sum_i \rho_i P_j}$ | NO       | NO        | NO         | NO        | NO      | Type I (can be extended to type II)         |
|                              | $P_{ij}>1$           |          |           |            |           |         |                                             |
| [65] (generalized formulation) | $g_{\text{max}} \frac{\phi_i P_j}{K + \sum_i \phi_i P_j}$ | NO       | NO        | NO         | NO        | NO      | Type II                                     |
|                              |                 |          |           |            |           |         |                                             |
| [63]                         | $P_{ij}(1+\beta P_j)$ | NO       | NO        | NO         | NO        | YES     | Type II for $K\beta<1$, otherwise type III |
|                              | $K + \sum_i \rho_i P_j$ |          |           |            |           |         |                                             |
| This paper, see also [71]    | $g_{\text{max}} \frac{\rho_i P_j}{\sum_i \rho_i P_j}$ | YES      | NO        | NO         | YES       | YES     | Types I, II, III (Depending on $F$)       |
|                              | $K + \sum_i \rho_i P_j$ |          |           |            |           |         |                                             |
| [23]                         | $g_{\text{max}} \frac{P_i}{K + \sum_i \rho_i P_j}$ | YES      | YES       | YES        | YES       | Supp. onstraints | Type II |
|                              | $K + \sum_i \rho_i P_j$ |          |           |            |           |         |                                             |
| This paper, Eqs.(12), (15–16)| $g_{\text{max}} \frac{\rho_i P_j}{\sum_i \rho_i P_j}$ | YES      | YES       | YES        | YES       | YES     | Types I, II, III (Depending on $F$)       |
| [11]                         | $g_{\text{max}} \frac{\rho_i P_j}{\sum_i \rho_i P_j}$ | YES      | YES       | YES        | Not applicable | YES    | Type II                                     |
|                              | $K + \sum_i \rho_i P_j$ |          |           |            |           |         |                                             |

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characterizes the rate of replenishment rate of nutrients (e.g. due to vertical diffusion in the water column).

A common sense expectation is that that the stationary total amount of prey $P$ and predator $P$ (determining ecosystem productivity) as obtained from (6)–(8) or any similar generic food chain model should not be sensitive with respect to the way we subdivide the prey species into subpopulations or groups. Indeed, such division can be purely conventional or even arbitrary (cf. the thought experiment above) and hence it should not determine the essential system properties. In particular, if we increase the number $n$ by splitting each prey species into to $m_i$ groups with the same life traits, this should not affect the total biomass of the given trophic level. Violation of Rule (iv) can result in an artificial bias in the relation between productivity and biodiversity which is predicted by some models; see examples in Section 5.

Rule (v). The total food intake rate should be an increasing function of densities of each of the resources. Mathematically this condition can be expressed as

$$\frac{\partial}{\partial P_j} \left( \sum_i \phi_i P_i \right) > 0. \quad (9)$$

Functional responses satisfying this condition are known in the literature as optimal functional responses. A functional response which does not satisfy Rule (v) is called suboptimal. Only an optimal functional response can have a solid ecological and evolutionary rationale [8], [11], [39]. Indeed, a suboptimal response effectively assumes that an extra food may act as a poison. A drop in the overall intake rate when a resource becomes more abundant can hardly have a solid theoretical justification and should be considered as a mathematical artifact [11], [39].

Rule (vi). Use of Holling type III to describe pairwise predator-prey interactions should be avoided unless it is carefully documented and justified.

For a given predator species, one should verify which type of response (according to the well-known Holling classification) takes place for feeding on a single food source. In particular, it has been argued [6], [39], [41] that a sigmoid functional response known as Holling type III may be a fingerprint of active switching: the intake of a given resource by the predator can drop quickly at small resource densities as a result of the predator switching to an alternative prey species, even if the latter is not considered explicitly. However, in the case where we explicitly describe all the possible resources in the system, the use of a sigmoid parameterization may not always be justified. A potential problem may occur in the case where the densities of all resources attain low values: since there are not any alternative resources, a simultaneous sharp drop in all $f_i(P)$ would then be a model artefact. Therefore, unless there is a clear argument supporting the use of Holling type III, Holling type II should be used.
Critical analysis of the existing formulations of the multi-prey functional response with switching

As a part of our critical analysis we have considered several multi-prey functional responses for a predator with food selectivity and switching which we found in the literature [8], [11–14], [23–26], [34], [60–65]. We have checked whether Rules (i)–(vi) are satisfied. Our results are summarized in Table 1 and details can be found in the online appendix; see Material S1. Rather surprisingly, none of the existing multi-prey functional responses with switching behavior satisfies all basic Rules (i)–(vi) indicating that those parameterizations may have some hidden caveats. In particular, most parameterizations (except for the one by [23] which we will analyse separately, see below) fail to account for the situation where species possess close life traits. Splitting a single population into groups and summing them back gives the initial functional response only in the absence of other resources; in particular, it does not work in the case of two distinctly different prey species.

Interestingly, the functional responses with proportion-based consumption ([11] see Table 1) satisfy all the basic rules (except Rule (iv) since in this case we do not have the possibility of coexistence of more than two species due to competitive exclusion) and hence are able to correctly describe consumption of prey species with similar traits. However, they fail to take into account switching.

Rule (vi) [Holling type for a single resource] is not satisfied for the parameterization suggested by [26]. Their parameterization is essentially of Holling type III and describes feeding of herbivorous zooplankton; however, this seems to be at odds with many field and laboratory studies. Indeed, a large number of laboratory experiments prove non-sigmoid Holling types I or II functional responses for zooplankton grazers [38], [66–68]. This is particularly true for the most efficient zooplankton grazers – microzooplankton-which usually exhibit non-sigmoidal Holling types I and II responses [69], [70]. Assuming a sigmoid model for zooplankton grazers can therefore result in a substantial underestimation of the intake rate at low overall biomasses of phytoplankton in models.

Recently, a new parameterization of the functional response was proposed in [23]. It was derived based on the biologically reasonable assumption that the predator possesses memory of the quality of the prey consumed previously, which defines its choice of the next prey item. Correspondingly, their parameterization seems to be more advanced compared to the previous attempts; indeed, as it is shown in Table 1, it satisfies Rules (i–iv) and (vi). However, optimality remains a problem: it is readily seen (see Material S1 for details) that Rule (v) does not hold unless some special constraints on the food similarity parameters $e_{ij}$ are introduced. In particular, for two distinctly different species (say, species 1 and species 2), the optimal intake is only possible for $e_{12}$ if $\max(e_{22}, e_{11})/2$. The applicability of the parameterization given in [23] is therefore restricted to communities with special properties.

Moreover, a biologically meaningful functional response with food selectivity should account for the food selectivity, which is a function of relative proportions of food densities [11] Mathematically, in case of a very low or a very high relative proportion of resources, it means that

$$\frac{f_i(P_1, P_2)}{f_j(P_1, P_2)} \to 0 \text{ for } P_i \to P_j, \; i \neq j$$

(11)

It is readily seen (see the online appendix for details) that, for the functional response proposed in [23], we obtain that, for $P_i/P_j \to 0$, $f_j/P_j \to e_{12}/\varepsilon_{ij}$. Apparently, it can only be consistent with food selectivity or switching if $e_{12} < \varepsilon_{ij}$ see Eq.(11). However, this clearly disagrees with the condition on the food similarity coefficients needed for optimality, see the previous paragraph, as the latter implies that $e_{12}/\varepsilon_{ij} > 1/2$. Therefore, for the functional response in [23], optimality and food selectivity appear to be mutually exclusive, at least in the range of large or small values of the prey population density. As a result, under the optimality condition, their parameterization of the functional response describes a proportionate resource consumption with no selectivity, except for relatively close values of $P_1$ and $P_2$.

The fact that there is no switching for small and large ratios of resource densities can have a major effect on model predictions. In particular, the per capita mortality due to predation on a species with very small density $P_i$ will remain non-zero as in the classical proportion-based consumption [11], [14] whereas in the case of (11) the rare species will be released from predation. At the same time, we want to mention it here that the fact that a functional response does not satisfy condition (11) does not necessarily mean that the corresponding parameterization is totally irrelevant or unrealistic. However, its switching properties do become restricted, especially in the important case where the density of one resource is much larger or smaller compared to the others. We therefore conclude that the problem of finding an appropriate parameterization for the predator multispecies response remains open. We address it in the next section.

Transitional multi-species functional response combining switching with proportion-based feeding

We shall now make an attempt to derive a universal, transitional functional response combining the features of both active and passive feeding. In doing this, we also aim to demonstrate that it is possible to construct a multispecies response that satisfies all Rules (i–vi), i.e. the constraints given by the rules are not mutually exclusive. As a starting point, we use the generic family of multi-prey responses as given in [11], [13], which has the following form:

$$f_i(P) = z_{i}(\bar{P})F \left( \sum_{i} \rho_{i} P_{i} \right).$$

(12)

where $z_{i}$ have the meaning of food preferences (which can be either constant or certain functions of $\bar{P}$, $\sum z_{i} = 1$, and $P$ is a function which describes the total consumption of all resources (e.g. Holling type II); $\rho_{i}$ are positive parameters (having the meaning of weights, however the sum of $\rho_{i}$ can be different from 1).

Depending on the choice of $z$, response (12) can attain different properties. In particular, it describes the proportion-based consumption if we choose

$$z_{i}(\bar{P}) = \frac{\rho_{i} P_{i}}{\sum_{i} \rho_{i} P_{i}}$$

(13)

(cf. [11]), and it can allow for switching if we choose $z_{i}$ differently. For instance, it is readily seen that the functional response with switching used in [71] (which is essentially a generalization of the previous results by [11] and [24]) coincides with (12) if we choose

$$z_{i}(\bar{P}) = \frac{\rho_{i} P_{i}}{\sum_{i} \rho_{i} P_{i}} \gamma > 1$$

(14)
Here we suggest the following parameterization of \( x_i \)
\[
x_i(\vec{P}) = \frac{P_i \rho_i \sum_j \varepsilon_{ij} P_j}{\sum_i \rho_i P_i \sum_j \varepsilon_{ij} P_j},
\]
(15)
where \( \varepsilon_{ij} \) are coefficients describing similarities between different types of food.

As we have argued above, for similar resources, i.e. for prey species with close life traits, consumption should be proportion-based (13), whereas for substantially distinct food sources the preference should be described by selectivity. In order to quantify closeness of different food sources in terms of the ability of the predator to distinguish between those sources, we introduce a certain parameter \( \omega_i \). This parameter describes a relevant species trait, for instance, the body size of prey, the defensive ability of prey, the body shape, mobility of organism, etc. (In a more general case \( \omega_i \) can be a vector taking into account many different traits.) Thus, similar types of food sources \( i \) and \( j \) should have close values \( \omega_i \) and \( \omega_j \) whereas distinct resources should have substantially different values.

Our main hypothesis here is that the food similarity coefficients are functions of their closeness so that for close prey species \( \varepsilon_{ij} \gg 0 \) and \( \varepsilon_{ij} \approx 0 \) for distinctly different prey species. For the sake of simplicity and to lessen the number of parameters, we consider that \( \omega_i = \rho_i \), thus assuming that similarity between food sources is sufficiently taken into account by the weights \( \rho_i \). For the functions \( \varepsilon_{ij} \) one can use, for instance, the Gaussian distribution.

\[
\varepsilon_{ij} = \varepsilon_0 \exp \left( -\frac{||\omega_i - \omega_j||^2}{2\sigma^2} \right) = \varepsilon_0 \exp \left( -\frac{(\rho_i - \rho_j)^2}{2\sigma^2} \right),
\]
(16)
where \( \varepsilon_0 \) is the normalizing coefficient (because the sum of \( \varepsilon_{ij} \) over \( j \) should be equal to unity) and \( \sigma \) is a parameter which depends on the sensitivity of the predator.

Now we demonstrate how the transitional parameterization (12), (15) and (16) works for the two important limiting cases.

Case A: all prey species are similar. Different prey species are not distinguished by the predator. In this case, the kernel (16) approaches the Dirac delta function centered at \( \rho_i \), and summation \( \sum_j \varepsilon_{ij} P_j \) give \( P_i \) so that (12), (15–16) turns into
\[
f_i(\vec{P}) = g_{\max} \frac{\rho_i P_i}{\sum_i \rho_i P_i} F \left( \sum_i \rho_i P_i \right)
\]
which clearly describes a response with switching. Thus, we have prey switching by the predator in the case of feeding on distinctly different resources.

Finally, it is easy to check that the functional response (12), (15–16) satisfies all the basic rules introduced in Section 2. Rules (i) and (ii) are satisfied by construction of \( f_i(\vec{P}) \); Rule (iii) is satisfied because for close species the consumption is described by the proportion-based formulation (13), for which this rule holds. The other rules are also satisfied, in particular, the a universal, transitional functional response is suggested optimal (in the sense of Rule (vi)) since the overall summation of (12) will give the resultant function \( F \) which we assume to be an increasing function of the food density.

Discussion

The choice of parameterization for the predator functional response is a key issue for modelling food web dynamics. It has attracted considerable attention recently, in particular because of the attempts to relate patterns of biodiversity in complex communities with ecosystem productivity/biomass [12], [44], [45], [72]. In a recent simulation study, Prowe et al. (see [12], [14]) showed the existence of a strongly positive correlation between the species richness and the biomass of primary producers in ecosystems with top-up control. In [14] it was found that the total biomass of primary producers \( \sum_i P_i \) in complex multi-species planktonic ecosystems exhibits a rapid increase in response to an increase in the species richness \( n \). This was observed both in a non-spatial and 3-D spatial models (see fig 4 and fig 8 in the cited paper). For example, transition from 4 to 20 primary producers would result in an increase in the total biomass approximately by a factor of 3.5. Moreover, a severe drop in the nutrient level was reported as a consequence of an increase in the system richness. However, a close inspection of the predator functional response used in [14] shows that it is not stable with respect to species subdivision; in particular, Rule (iv) above is not satisfied. A several-fold time increase in the total biomass could be observed if a single phytoplankton species is split into several groups that are indistinguishable to consumers (see the description of the thought experiment in Section 2). Our analysis based on the basic food chain model (6)–(8) (see Material S1) can easily explain the findings in [14]. Indeed, one can see that \( \sum_i P_i \propto n \) within a certain range of \( n \) (see [81] and [89] in Material S1). Analysis of the expression for the level of the limiting nutrient \( N \) shows (see [S8] in Material S1) that it is a decreasing function of species richness \( n \). Correspondingly, one can observe a similar increase in biodiversity and transmission to oligotrophic conditions (a severe drop in \( N \)) not by increasing the species richness but by just splitting a single phytoplankton species into \( n \) groups. It indicates that the findings [14] can be specific to their choice of the functional response, effectively being a model artefact.

Therefore, one should be very careful when choosing a parameterization of the functional response of a generalist predator in multi-species models as it may affect the model predictions. We should mention here that, even in the simple case where the predator consumes only one type of resource, variation...
in the mathematical formulation of the functional response can affect the model properties significantly [73–75], a phenomenon known as structural sensitivity of biological models [75], [76]. Considering predation on multiple resources makes food web models even more sensitive to the choice of functional response parameterization [13], [8], [64]. However, unlike the case of a single-prey functional response, the conventional way to choose the parameterization of a multi-prey functional response is mostly theoretical [11], [59], mainly due to limited empirical data.

In this paper, we endeavour to make the process of selection of the functional response more rational by introducing certain basic rules (see Section 2) specifically designed in order to avoid model artefacts. In particular, we argue that, in truly multi-species systems (such as plankton communities [46]; [48]), insect communities [51] or coral reef ecosystems [52], one can hardly reduce the biological rationale behind the functional response to the case of clearly different resources allowing for food selectivity of the predator. In the cases where a consumer has to choose between hundreds or even thousands of different food sources, the predator will often be unable to distinguish between prey species with close life traits. Hence it will show switching only between distinct food sources but has to employ proportion-based feeding on close food sources. Surprisingly, this argument seems to have been somewhat ignored in the derivation of the most of the recently proposed functional responses (but see [23]).

We propose a novel theoretical framework for construction of a transitional multi-prey functional response (12), (15) and (16) which combines food selectivity and switching with proportion-based consumption. We quantify the entire spectrum of different resources according to their relative closeness to each other, i.e. food resources with close types are (virtually) undistinguishable by the predator. Introducing similarity of the resources has an advantage over the previous formulations as it allows us to combine switching for distinct types of food with proportion-based consumption for similar types into a single mathematical expression. The new multi-prey functional response (12), (15) and (16) satisfies Rules (i)–(vi). Note that our approach can be readily extended to the case of a continuous spectrum of resources (as can arise, for instance, when one takes into account an inherently continuous statistical variation of individual traits, cf. [77]).

We mention here that the rules introduced in this paper (see Section 2) are necessary but by no means exhaustive. Particularities of a given ecosystem and of specific prey and/or predator species may impose supplementary requirements and constraints on the choice of the functional response parameterization. For instance, any interference between the predators (e.g. [54], [64]) is likely to affect their prey selectivity. Also, the existing parameterizations of the predator functional response mostly focus on the resource densities; however, the handling times are likely to have a similar effect on the predator’s choice of food. Indeed, an optimal strategy for a predator would likely be to consume the resources with a small handling time and to avoid those with large handling times [78], especially if the predator has a sufficiently long memory (cf. [23]) and the resources are abundant, so food handling would basically determine the entire intake rate. It seems reasonable to expect that the predator then should exhibit switching, at least when it feeds on a few resources that are sufficiently different in terms of their handling time. Surprisingly, it is readily seen that none of the functional responses listed in Table 1 is capable to describe switching with respect to the handling time, all of them showing only proportion-based prey consumption. A better understanding of this issue should become a focus of future research.

### Supporting Information

**Material S1** Verification of the Basic Rules for the multi-prey functional responses listed in Table 1. (DOC)

**Acknowledgments**

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**Author Contributions**

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### References

1. Steinke M, Malin G, Lis P (2002) Trophic interactions in the sea: An ecological role for climate relevant volatiles? J Phycol 38: 630–638.
2. Cohen JE, Jonsson T, Carpenter SR (2003) Ecological community description using the food web, species abundance, and body size. Proc Natl Acad Sci USA 100: 1781–1786.
3. Odum EP, Barrett GW (2005) Fundamentals of ecology. Thompson Brooks/ Cole, Belmont, CA.
4. Begon M, Townsend CR, Harper JL (2005) Ecology: from individuals to ecosystems, 6th Edition. Michael. 752p.
5. Masłowski M, Gravel D, Mouquet N, Cadotte MW, Fukami T, et al. (2011) Linking community and ecosystem dynamics through spatial ecology. Ecol Lett 14: 313–323.
6. Murdoch WW, Oaten A (1975) Predation and population stability. Adv Ecol Res 9: 1–131.
7. Comins HN, Hassell MP (1976) Predation in multi-prey communities. J Theor Biol 52: 93–114.
8. Holt RD (1983) Optimal foraging and the form of the predator isolcline. Amer Nat 122: 321–341.
9. Armstrong RA (1999) Stable model structures for representing biogeographical diversity and size spectra for plankton communities. J Plankton Res 21: 445–64.
10. Krivan V, Slatker A (1999). Optimal foraging and predator-prey dynamics II. Theor Popul Biol 5: 111–126.
11. Gentleman W, Leising A, Frost B, Storm S, Murray J (2003). Functional response parameterization [13], [8], [64]). However, unlike the case of a single-prey functional response, the conventional way to choose the parameterization of a multi-prey functional response is mostly theoretical [11], [59], mainly due to limited empirical data. Considering predation on multiple resources makes food web models even more sensitive to the choice of functional response parameterization [13], [8], [64]. However, unlike the case of a single-prey functional response, the conventional way to choose the parameterization of a multi-prey functional response is mostly theoretical [11], [59], mainly due to limited empirical data.

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1. Steinke M, Malin G, Lis P (2002) Trophic interactions in the sea: An ecological role for climate relevant volatiles? J Phycol 38: 630–638.
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3. Odum EP, Barrett GW (2005) Fundamentals of ecology. Thompson Brooks/ Cole, Belmont, CA.
4. Begon M, Townsend CR, Harper JL (2005) Ecology: from individuals to ecosystems, 6th Edition. Michael. 752p.
5. Masłowski M, Gravel D, Mouquet N, Cadotte MW, Fukami T, et al. (2011) Linking community and ecosystem dynamics through spatial ecology. Ecol Lett 14: 313–323.
6. Murdoch WW, Oaten A (1975) Predation and population stability. Adv Ecol Res 9: 1–131.
7. Comins HN, Hassell MP (1976) Predation in multi-prey communities. J Theor Biol 52: 93–114.
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9. Armstrong RA (1999) Stable model structures for representing biogeographical diversity and size spectra for plankton communities. J Plankton Res 21: 445–64.
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11. Gentleman W, Leising A, Frost B, Storm S, Murray J (2003). Functional response parameterization [13], [8], [64]). However, unlike the case of a single-prey functional response, the conventional way to choose the parameterization of a multi-prey functional response is mostly theoretical [11], [59], mainly due to limited empirical data. Considering predation on multiple resources makes food web models even more sensitive to the choice of functional response parameterization [13], [8], [64]. However, unlike the case of a single-prey functional response, the conventional way to choose the parameterization of a multi-prey functional response is mostly theoretical [11], [59], mainly due to limited empirical data.
26. Ryabchenko VA, Fasham MJR, Kagan BA, Popova EE (1997). What causes short-term oscillations in ecosystem models of the ocean mixed layer? J Marine Syst 13: 53–50.

27. Stoeker DK, Cucci TL, Hultbart EM, Yentsch CM (1986) Selective feeding by Balanus sp. (Ciliata: Balanidae) on Phaeocystis that best support its growth. J Exp Mar Biol Ecol 95: 113–130.

28. Colton TF (1987) Extending functional response models to include a second prey type: an experimental test. Ecology 68: 900–912.

29. Kiørboe T, Saiz E, Viitasalo M (1996). Prey switching behaviour in the planktonic copepod Acartia tonsa. Mar Ecol Prog Ser 143: 63–73.

30. Strom SL, Loukos H (1998) Selective feeding by protozoa: model and experimental behaviors and their consequences for population stability. J Plankton Res 20: 831–846.

31. Leising AW, Pierton JJ, Halsband-Leuk C, Horner RA, Postel J (2005) Copepod grazing during spring blooms: Does Calanus finmarchicus avoid harmful diatoms? Prog Oceanogr 67: 384–405.

32. Nejstgaard JC, Tang KW, Steinke M, Dutz J, Koski M, et al (2007) Zooplankton grazing on Phaeocystis: a quantitative review and future challenges. Biogeochemistry 83: 147–172.

33. Elliott JM (2006). Prey switching in Rhaphidophora denisae (tri chlorophyta) alters with larval instar. Freshwater Biol 51: 913–924.

34. Stoner MC, Ekerberg CJ, Matiasopoulos C, Fernández S, Redpath S, et al (2010) The functional response of a generalist predator. PloS One 5(5): e10761.

35. Diekmann O, Gyllenberg M, Metz JA, Nakaoka S, de Roos AM (2010) Daphnia revisited: local stability and bifurcation theory for phytoplankton structured population models explained by way of an example. J Math Biol 61: 277–318.

36. Johnson PTJ, Stanton DE, Preu ER, Forshay KJ, Carpenter SR (2006) Dining on disease: how interactions between infection and environment affect predation risk. Ecology 87: 1973–1980.

37. Frankis PJS (2009) Planktonic ecosystem models: perplexing parameterizations and a failure to fail. J Plankton Res 31: 1299–1306.

38. Naeem S, Li S (1997) Biodiversity enhances ecosystem reliability. Nature 390: 66–74.

39. Vallina SM, Ward BA, Dutkiewicz S, Follows MJ (2012) Maximal foraging with a failure to fail. J Plankton Res 31: 1299–1306.

40. Haydon DT (1994) Pivotal assumptions determining the relationship between competitive preys. J Theor Biol 79: 303–315.

41. Vance RR (1978) Predation and resource partitioning in one-predator–two-prey model communities. Amer Nat 112: 797–813.

42. Evans GT, Garçon VC (ed) (1997) One-dimensional models of water column biogeochemistry. Joint Global Ocean Flux Study Rep. No. 25/97. JGOFS, Bergen, Norway.

43. Koen-Mono M, Yodzis P (2005). Multi-species modelling of some components of the marine community of northern and central Patagonia, Argentina. J Fish Aquat Sci 62: 1490–1512.

44. Koen-Mono M, Yodzis P (2005). Multi-species modelling of some components of the marine community of northern and central Patagonia, Argentina. J Fish Aquat Sci 62: 1490–1512.

45. Huston M (1985) Variation in coral growth rates with depth at Discovery Bay, Jamaica. Coral Reefs 4: 19–25.

46. Pace ML, Glasser JE, Pomeroy LR (1984) A simulation analysis of continental shelf food webs. Mar Biol 8: 47–63.

47. Raymont JEG (1980) Plankton and Productivity in the Oceans. Phytoplankton, marine zooplankton. Chemical cycles in the sea. Plenum Press, New York, 550p.

48. Edwards AM, Brindley J (1961) The paradox of the plankton. Amer Nat 95: 137–145.

49. Johnson PTJ, Stanton DE, Preu ER, Forshay KJ, Carpenter SR (2006) Dining on disease: how interactions between infection and environment affect predation risk. Ecology 87: 1973–1980.

50. Venter JC, Remington K, Heidelberg JF, Halpern AL, Rusch D., et al (2004) The complete genome of two complete microbial communities. Science 304: 66–74.

51. Erwin TL (1982) Tropical forests: their richness in Coleoptera and other arthropod species. Coleopterists Bull 36: 74–75.

52. Huston M (1985) Variation in coral growth rates with depth at Discovery Bay, Jamaica. Coral Reefs 4: 19–25.

53. Arditi R, Ginzburg LR (1989) Coupling in predator-prey dynamics: ratio-dependence. J Theor Biol 139: 311–326.

54. Abrams P and Ginzburg LR (2000) The nature of predation: prey dependent, ratio dependent or neither? TREE 15: 337–341.

55. Morozov AY, Arashkevich AG (2012) Towards a correct description of zooplankton feeding in models: Taking into account food-mediated unsynchronized vertical migration. J Theor Biol 262: 346–360.

56. Franks PJS (2009) Planktonic ecosystem models: perplexing parameterizations and a failure to fail. J Plankton Res 31: 1299–1306.

57. Edwards AM, Brindley J (1999) Zooplankton mortality and the dynamical behavior of plankton population models. Bull Math Biol 61: 202–339.

58.置于 A, Fiksen Ø (2013) Optimal foraging in marine ecosystem models: selectivity, profitability and switching. Mar Ecol Prog Ser 478: 91–101.