Clinicopathological characteristics and prognostic value of the cancer stem cell marker ALDH1 in ovarian cancer: a meta-analysis

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Background: The clinicopathological and prognostic values of the cancer stem cell marker aldehyde dehydrogenase 1 (ALDH1) in ovarian cancer (OC) remain unknown. The aim of our meta-analysis was to evaluate ALDH1’s association with clinicopathological characteristics and its prognostic significance in patients with OC.

Materials and methods: PubMed, Embase, and China Biology Medicine were systematically searched for eligible studies (up to October 2017). Pooled odds ratios (ORs) or hazard ratios (HRS) with 95% CIs were used to evaluate the association of ALDH1 expression with clinicopathological features and survival outcomes.

Results: A total of 17 papers (18 studies) that included 2,531 patients with OC were analyzed. The results showed a significant association between increasing ALDH1 expression and International Federation of Gynecology and Obstetrics stage (OR 2.02, 95% CI 1.16–3.52), lymph node metastasis (OR 1.91, 95% CI 1.01–3.61), and distant metastasis (OR 5.43, 95% CI 1.44–20.42) in OC. However, no significant correlation was found between increasing ALDH1 expression and age (OR 0.90, 95% CI 0.25–3.28), tumor size (OR 1.13, 95% CI 0.75–1.71), tumor location (OR 0.69, 95% CI 0.22–2.13), ascite status (OR 0.74, 95% CI 0.49–1.11), resistance status (OR 0.70, 95% CI 0.14–3.51), or clinicopathological type (OR 1.14, 95% CI 0.69–1.86). Moreover, a high ALDH1 expression was significantly associated with overall survival (HR 1.56, 95% CI 1.21–2.02) but not with disease-free survival (HR 1.38, 95% CI 0.99–1.93).

Conclusion: The meta-analysis indicates that increasing ALDH1 predicts poor prognosis and clinicopathological characteristics in OC. Future studies are needed to explore tailored treatments that directly target ALDH1 for the improvement of survival in OC.

Keywords: ovarian cancer, ALDH1, prognosis, clinicopathological characteristics, meta-analysis

Introduction

Ovarian cancer (OC) is the most common lethal gynecological malignancy, followed by cervical cancer and endometrial cancer, and has the third highest morbidity rate of gynecological malignancies. Based on its mortality rate, OC is first among all gynecologic malignancies.1 OC is characterized by its insidious onset, aggressiveness, rapid metastasis, and poor prognosis. Furthermore, given the lack of specific symptoms and methods for early screening, many patients with epithelial ovarian cancer (EOC) are diagnosed at an advanced stage, and their 5-year survival rate is only 25%–30%.2 In the clinic, EOC is the most common histological subtype of OC, which also includes serous OC, clear cell OC, endometrioid adenocarcinoma, and mixed cell OC. Several independent prognostic factors such as
disease stage, age, metastasis, classification, and residual tumor bulk have been identified in previous investigations.

Cancer stem cells (CSCs) are defined as a small population of cancer cells that have high tumorigenicity. Furthermore, the CSC hypothesis has been formulated to explain tumor occurrence and recurrence. Increasing evidence has indicated that CSCs have self-renewal capacity; therefore, CSCs are thought to be responsible for tumor progression, metastasis, and therapeutic resistance.

Aldehyde dehydrogenase 1 (ALDH1), an important marker of stem-like cells in many malignant tumors, is a zinc-containing enzyme with oxidation and detoxification functions. To this day, the correlation between ALDH1 expression and OC has been widely studied, but the results remain controversial. Several studies have suggested that the high expression of ALDH1 is related to disease stage, age, metastasis, classification, and residual tumor bulk, but other studies did not reach these conclusions. We thus conducted a meta-analysis to reveal the prognostic significance of ALDH1 for overall survival (OS) and disease-free survival (DFS) as well as the association between ALDH1 and clinicopathological features in patients with OC.

Materials and methods

Literature search

Several electronic databases, including PubMed, Embase, and China Biology Medicine (CBM), were independently searched by two researchers (W Zhao and C Zang) through October, 2017. We used the MeSH/Emtree terms, combining free text words that were properly adjusted for the different databases in all of the search strategies. The following keywords were used for searching: (“ALDH1” OR “aldehyde dehydrogenase 1” OR “ALDH1 enzyme”) AND (“ovarian cancer” OR “ovarian tumor” OR “ovarian carcinoma” OR “ovarian neoplasm”). For a detailed search strategy please refer to the Supplementary materials.

Inclusion and exclusion criteria

All of the studies in this meta-analysis had to meet the following inclusion criteria: 1) patient diagnoses were confirmed as OC by pathology, 2) the outcome of interest was the clinicopathological and/or prognostic relationship between ALDH1 and OC, 3) immunohistochemistry (IHC) detection methods were used to detect specific ALDH1 antigens with monoclonal anti-human ALDH1, and 4) the outcome measures of interest could be extracted directly or calculated from the published data indirectly. The exclusion criteria were as follows: 1) studies that were not performed in humans, and 2) studies that were not performed in humans, and 3) papers that failed to report sufficient data.

Data extraction

Extraction of data was independently performed by two reviewers (W Zhao and C Zang) in accordance with a predefined and standardized form, and discrepancies were settled through a consensus discussion. The following information was extracted from the included studies: first author, year, patient source, type of patients, median/mean age, sample size, technique, cutoff value, median follow-up, OS/DFS, and clinicopathological features. For the specific data extracted, see refer to the Supplementary materials.

Quality assessment

The quality of the included cohort studies was evaluated according to the Newcastle–Ottawa Scale (NOS). Two reviewers (W Zhao and C Zang) evaluated all of the studies independently, and discrepancies were solved by consensus. The standards included three aspects: 1) selection of the research groups, 2) comparability of the groups, and 3) measurement of exposure factors or outcome. Scores ranged from 0 to 9, and a score ≥6 indicated a high quality.

Statistical analysis

Data for the clinicopathological features were extracted from available studies of odds ratios (ORs). The hazard ratios (HRs) and 95% CIs were used to summarize survival outcomes. We would have obtained pooled HRs and 95% CIs provided the statistical data were reported in the study. When HRs and 95% CIs were not directly reported in the studies, survival information was extracted from the Kaplan–Meier curves using Engauge Digitizer 4.1 and used to estimate HR. Statistical heterogeneity was assessed using both a chi-square-based Q-test and I² statistic, which describe the percentage of total variation across studies caused by heterogeneity rather than by chance. Studies with an I² >50% or P<0.1 were considered to have significant heterogeneity, and a random-effects model test was conducted. Otherwise, the fixed-effects model test was adopted. Funnel plots and Egger’s test were used to evaluate publication bias. Sensitivity analysis was carried out to evaluate the effect of a single study on all the results and to find the origin of heterogeneity. P<0.05 was considered to be statistically significant. All the calculations were performed by Stata 13.0 (StataCorp LP, College Station, TX, USA).
Results

Literature search

There were 165 records found from database searches according to our search strategy, of which 45 papers were duplicates. Through screening title and abstract, 24 irrelevant papers were excluded; then 79 papers, including 13 reviews, 33 experimental studies, 28 papers lacking of data, 3 papers without immunohistochemistry (IHC) of ALDH1 detection and 2 papers unable to be obtained, were excluded after reading the full text. Finally, 17 eligible papers (18 studies) involving 2,531 patients were included in our meta-analysis. The detailed literature search and paper selection are shown in Figure 1.

Characteristics of the included papers

The details of the 17 included manuscripts, including 2,531 patients (sample size median: 88 [37–442]), selected from the literature search are summarized in Table 1. In total, 13 eligible papers (14 studies) with 1,841 patients were analyzed for clinicopathological features, and 15 qualified studies with 2,379 patients were analyzed for survival outcomes. The studies were conducted in seven countries (China, Finland, Korea, Japan, Germany, Norway, and the USA) and were published between 2009 and 2017. All included studies detected ALDH1 using IHC. Among included papers, ALDH1 expression in tumor cells was most commonly evaluated based on the percentage of positively stained cells or the semi-quantitative immunoreactivity score.

Quality assessment

Each of the 18 included studies was evaluated using the NOS, as described previously. Scores are shown in Table 2, in which “★” represents 1 point, “×” represents 0 point, and “—” represents uncertain points. Deng’s quality evaluation was not carried out due to lack of detailed information on the exposed group and the nonexposed group. Except for studies of Xue et al and Jing et al, the others are of high quality. Owing to lack of follow-up, the studies of Xue et al and Jing et al only have 5 points. Research of Kuroda et al and Wang et al could not score because follow-up occurred before the study, nor could studies of Chang et al and Huang et al, because they failed to control comparisons among groups. The studies of Ayub et al, Liebscher et al, and Landen et al did not describe the inter-group comparability. However, studies of Kuroda et al and Ruscito et al strictly controlled the inter-group comparability. The remaining studies only had one point in inter-group comparability. The differences between the inter-groups may primarily stem from age, tumor stage, grade, or lymph node metastasis.

Effect of ALDH1 expression on clinicopathological parameters

Age, tumor size, tumor location, ascite status, resistance status, and clinicopathological type

We assessed the association between age, tumor size, tumor location, ascite status, resistance status, clinicopathological type, and ALDH1 expression. As illustrated in Table 3, ALDH1 expression was not associated with age (<55 vs ≥55: OR 0.90, 95% CI 0.25–3.28, $I^2$ 74.1%), tumor size (<8 cm vs ≥8 cm: OR 1.13, 95% CI 0.75–1.71, $I^2$ 0.0%), tumor location (unilateral vs bilateral: OR 0.69, 95% CI 0.22–2.13, $I^2$ 71.5%), ascite status (yes vs no: OR 0.74, 95% CI 0.49–1.11, $I^2$ 0.0%), resistance status (yes vs no: OR 0.70, 95% CI 0.14–3.51, $I^2$ 87.6%), or clinicopathological type (S vs C: OR 1.14, 95% CI 0.69–1.86, $I^2$ 0.0%). The results are shown in Figure 2A–F and Table 3.

International Federation of Gynecology and Obstetrics (FIGO) stage

The incidence of ALDH1 in FIGO stage may indicate the degree of FIGO stage (III–IV vs I–II: OR 2.02, 95% CI 1.16–3.52, $I^2$ 78.0%). Subgroup analyses were adopted on sample size (<100: OR 3.34, 95% CI 2.20–5.08,
Table 1 Characteristics of the studies included in the present meta-analysis

| Studies          | Origin of population | Quality scores | n    | Age (years) | Cutoff value                      | Follow-up time (months) | FIGO stage | Histological type | Survival analysis | HR |
|------------------|----------------------|----------------|------|-------------|----------------------------------|-------------------------|-------------|-------------------|-------------------|----|
| Kuroda et al     | Japan                | 8              | 62   | 56.6±1.4    | Extent 20% (high/low)            | 60                      | I–IV        | S                 | OS                | SC |
| Chang et al      | USA                  | 7              | 37   | 52.2±9.4    | Extent 15% (High/Low)            | 60                      | C           |                  |                   |    |
| Chang et al      | USA                  | 7              | 442  | 60 (21–89)  | Extent 20% (high/low)            | 96.0 (78.0–122.4)       | I–IV        | Mixed             | OS                | SC |
| Ayub et al       | Germany              | 7              | 55   | 55±8.9      | IRS (high/low)                   | 44.81 (17.01–72.61)     | IIIC–IV     | S                 | OS                | Reported |
| Ayub et al       | Germany              | 7              | 131  | 59.6 (34.0–87.1) | IRS≥4 (negative/positive)     | 200                     | I–IV        | S                 | OS                | Reported |
| Liebscher et al  | Germany              | 7              | 112  | 56 (33–77)  | IRS≥2 (negative/positive)       | 101 (4–475)             | I–IV        | S                 | OS                | SC |
| Chen et al       | China                | 8              | 80   | NR          | Extent 10% (negative/positive)   | 144                     | I–IV        | Mixed             | OS                | SC |
| Yu et al         | China                | 8              | 207  | 59.1 (22–75) | IRS≥3 (negative/positive)      | <100                    | I–IV        | Mixed             | OS                | Reported |
| Ruscito et al    | European             | 9              | 112  | 53.0±2.8    | IRS (negative/positive)         | <250                    | I–IV        | S                 | OS                | SC |
| Sun et al        | China                | 8              | 100  | NR          | Extent 10% (high/low)           | 150–200                 | I–IV        | C                 | OS                | SC |
| Mizuno et al     | Japan                | 8              | 81   | 3 (31–82)   | Extent 10% (high/low)           | 150–200                 | I–IV        | C                 | OS                | SC |
| Huang et al      | Norwegian            | 7              | 248  | 58 (19–89)  | IRS (negative, low, high)       | 300–360                 | I–IV        | Mixed             | OS                | Reported |
| Landen et al     | USA                  | 7              | 65   | 62.2 (34–89) | Extent 1% (high/low)            | >100                    | III–IV      | S                 | OS                | SC |
| Kim et al        | Korea                | 8              | 38   | 46.8 (33–83) | Extent (high/low≥2-fold)        | 32 (9–86)               | I–IV        | C                 | OS                | Reported |
| Yu et al         | China                | 8              | 198  | 51.7 (19–75) | IRS≥3 (negative/positive)      | 6–110                   | I–IV        | Mixed             | OS                | Reported |
| Xue et al        | China                | 5              | 60   | 53.6 (42–70) | IRS≥3 (negative/positive)      | NR                      | I–IV        | Mixed             | NR                | NR |
| Jing et al       | China                | 5              | 92   | 51 (19–80)  | IRS≥4 (negative/positive)       | NR                      | I–IV        | Mixed             | NR                | NR |

Note: Extent, the numbers of stained cells.

Abbreviations: FIGO, International Federation of Gynecology and Obstetrics; HR, hazard ratio; S, serous ovarian carcinoma; C, clear cell adenocarcinoma; OS, overall survival; DFS, disease-free survival; SC, survival curve; IRS, immunoreactivity score; NR, not reported.
Table 2 Quality scores of included studies using NOS

| Studies               | Representativeness of exposed group | Representativeness of non-exposed group | Sources of exposure factor | Not observed outcome at first | Comparability | Evaluation of outcome | Enough follow-up time | Adequacy | Scores |
|-----------------------|-------------------------------------|----------------------------------------|---------------------------|----------------------------|---------------|-----------------------|-----------------------|-----------|--------|
| Kuroda et al          | ★                                   | ★                                      |                           | ★                         | ★             | ★                     | ★                     |           | 8      |
| Chang et al           | ★                                   | ★                                      | ★                         | ×                         | ★             | ★                     | ★                     |           | 7      |
| Ayub et al            | ★                                   | ★                                      | ★                         | ×                         | ★             | ★                     | ★                     |           | 7      |
| Wang et al            | ★                                   | ★                                      | ★                         | ×                         | ★             | ★                     | ★                     |           | 7      |
| Liesbacher et al      | ★                                   | ★                                      | ★                         | ×                         | ★             | ★                     | ★                     |           | 7      |
| Deng et al            | ★                                   | ★                                      | ★                         | ×                         | ★             | ★                     | ★                     |           | 7      |
| Chen et al            | ★                                   | ★                                      | ★                         | ×                         | ★             | ★                     | ★                     |           | 8      |
| Yu et al              | ★                                   | ★                                      | ★                         | ×                         | ★             | ★                     | ★                     |           | 8      |
| Ruscito et al         | ★                                   | ★                                      | ★                         | ×                         | ★             | ★                     | ★                     |           | 9      |
| Mizuno et al          | ★                                   | ★                                      | ★                         | ×                         | ★             | ★                     | ★                     |           | 8      |
| Huang et al           | ★                                   | ★                                      | ★                         | ×                         | ★             | ★                     | ★                     |           | 7      |
| Landen et al          | ★                                   | ★                                      | ★                         | ×                         | ★             | ★                     | ★                     |           | 7      |
| et al                 | ★                                   | ★                                      | ★                         | ×                         | ★             | ★                     | ★                     |           | 5      |

Notes: "★" represents 1 point; "×" represents 0 point; and "—" represents uncertain points.

Abbreviation: NOS, Newcastle–Ottawa Scale.

40.1%; ≥100: OR 1.18, 95% CI 0.53–2.60, I² 85.1%) and ethnicity (Asian: OR 2.82, 95% CI 2.13–3.72, I² 23.1%; non-Asian: OR 0.45, 95% CI 0.29–0.71, I² 0.0%). The results are shown in Figure 3A and Table 3.

Histopathologic grade

Pooled OR from eight studies is shown in Figure 3B and Table 3, indicating that high ALDH1 expression was not associated with the degree of histopathologic grade (G1 vs G2+G3: OR 0.53, 95% CI 0.23–1.27, I² 83.2%). Subgroup analyses were only adopted on sample size (<100: OR 0.29, 95% CI 0.06–1.38, I² 81.9%; ≥100: OR 0.87, 95% CI 0.30–2.54, I² 85.2%) due to the limited number of studies on ethnicity.

Lymph node metastasis

The pooled OR from eight studies is shown in Figure 3C and Table 3, indicating that a high ALDH1 expression was

Table 3 Effects of ALDH1 expression on clinicopathological parameters and survival outcome

| Clinicopathological variables | No of studies | Cases | Selected model | Heterogeneity test | OR/HR (95% CI) | Z  | P²  | P<sub>Begg</sub> |
|------------------------------|---------------|-------|----------------|--------------------|----------------|----|-----|------------------|
| Age (>50 years/<50 years)    | 2             | 329   | Random         | 74.1               | 0.049          | 0.90 (0.25–3.28) | 0.16 | 0.870 –          |
| Tumor size (≥8 cm/<8 cm)     | 2             | 405   | Fixed          | 74.0               | 0.900          | 1.13 (0.75–1.71) | 0.58 | 0.562 –          |
| Tumor grade (G1/G2+G3)       | 8             | 1,067 | Random         | 83.2               | 0.000          | 0.53 (0.23–1.27) | 1.42 | 0.156 0.386     |
| Tumor stage (II+IV/I–II)     | 14            | 1,841 | Random         | 78.0               | 0.000          | 2.02 (1.16–3.52) | 2.50 | 0.012 0.743     |
| LNM (yes/no)                 | 8             | 1,176 | Random         | 67.9               | 0.003          | 1.91 (1.01–3.61) | 1.98 | 0.048 1.000     |
| Ascite (yes/no)              | 2             | 405   | Fixed          | 74.0               | 0.869          | 0.74 (0.49–1.11) | 1.45 | 0.147 –          |
| Chemoresponse (sensitive/resistant) | 3     | 592   | Random         | 87.6               | 0.000          | 0.70 (0.14–3.51) | 0.44 | 0.664 –          |
| Tumor location (unilateral/bilateral) | 2     | 287   | Random         | 71.5               | 0.061          | 0.69 (0.22–2.13) | 0.65 | 0.517 –          |
| Distant metastasis (yes/no)  | 2             | 245   | Random         | 58.6               | 0.120          | 5.43 (1.44–20.42) | 2.50 | 0.012 –          |
| Clinicopathological type (S/C) | 8       | 1,450 | Random         | 74.0               | 0.097          | 1.19 (0.69–1.86) | 0.50 | 0.614 0.711     |
| OS                           | 15            | 2,279 | Random         | 74.0               | 0.000          | 1.56 (1.21–2.02) | 3.39 | 0.001 0.083     |
| DFS                          | 10            | 1,477 | Random         | 76.1               | 0.000          | 1.38 (0.99–1.93) | 1.92 | 0.055 0.788     |

Abbreviations: ALDH1, aldehyde dehydrogenase 1; OR, odds ratio; HR, hazard ratio; CI, confidence interval; LNM, lymph node metastasis; S, serous ovarian carcinoma; C, clear cell adenocarcinoma; OS, overall survival; DFS, disease-free survival.
significantly associated with lymph node status (yes vs no: OR 1.91, 95% CI 1.01–3.61, $I^2$ 67.9%). Subgroup analyses were adopted on sample size (<100: OR 2.42, 95% CI 0.82–7.11, $I^2$ 67.6%; ≥100: OR 1.91, 95% CI 1.00–3.60, $I^2$ 75.0%). All patients on lymph node metastasis were from Asia.

**Distant metastasis**

The ORs for distant metastasis were only available in two studies, indicating that high ALDH1 expression was significantly associated with distant metastasis (yes vs no: OR 5.43, 95% CI 1.44–20.42, $I^2$ 58.6%). The results are shown in Figure 3D and Table 3.

**Effect of ALDH1 expression on survival outcome**

We assessed the association between increasing ALDH1 levels and survival outcomes of OC patients. In total, 10 eligible studies with 1,477 patients were analyzed for DFS, and 15 qualified studies with 2,279 patients were analyzed for OS. The results are shown in Figure 3E and F and Table 3, indicating that a high ALDH1 expression was significantly associated with OS (high [H] vs low [L]: HR 1.56, 95% CI 1.21–2.02, $I^2$ 74.0%) but not with DFS (H vs L: HR 1.38, 95% CI 0.99–1.93, $I^2$ 76.1%). Subgroup analyses for OS were adopted on sample size (<100: HR 2.37, 95% CI 1.79–3.14, $I^2$ 0.0%; ≥100: HR 1.16, 95% CI 0.92–1.46, $I^2$ 72.1%) and ethnicity (Asian: HR 2.07, 95% CI 1.64–2.63, $I^2$ 0.0%; non-Asian: HR 1.18, 95% CI 0.90–1.54, $I^2$ 68.5%). The subgroup analyses for DFS were performed on sample size (<100: HR 1.50, 95% CI 1.11–2.03, $I^2$ 46.3%; ≥100: HR 1.26, 95% CI 0.78–2.04, $I^2$ 87.2%) and ethnicity (Asian: HR 1.45, 95% CI 1.07–2.01, $I^2$ 55.2%; non-Asian: HR 1.34, 95% CI 0.86–2.07, $I^2$ 84.0%).

**Publication bias**

Owing to the limited number of OC studies, we did not assess publication bias using Beggs’s test, Egger’s test or
funnel plots for age, tumor size, tumor location, ascite status, and resistance. The funnel plots appeared to be symmetric, showing no evidence of publication bias for the other pooled clinicopathological characteristics and survival outcomes. In addition, both tests also indicated that there was no publication bias for pooled tumor grade ($P_{\text{Begg}} = 0.231$), FIGO stage ($P_{\text{Begg}} = 0.242$), lymph node metastasis ($P_{\text{Begg}} = 0.950$), or clinicopathological type ($P_{\text{Begg}} = 0.868$) as well as DFS.

Figure 3 Forest plot of ORs/HRs for the association between ALDH1 expression and FIGO stage, histopathologic grade, lymph node metastasis, distant metastasis, DFS, and OS. Notes: (A) FIGO stage (II–IV vs I–II), (B) histopathologic grade (G1 vs G2+G3), (C) lymph node metastasis (yes vs no), (D) distant metastasis (yes vs no), (E) DFS (H vs L), and (F) OS (H vs L). Weights are from random-effects analysis.

Abbreviations: ALDH1, aldehyde dehydrogenase 1; OR, odds ratio; HR, hazard ratio; FIGO, International Federation of Gynecology and Obstetrics; DFS, disease-free survival; OS, overall survival; H, high; L, low.
(P = 0.788) or OS (P = 0.083). The results are shown in Figure 4A–F and Table 3.

Sensitivity analysis
We performed a sensitivity analysis to assess the stability of our results regarding clinicopathological characteristics, DFS, and OS in OC patients, and results are shown in Figure 5A–F. No significant changes were identified in the pooled results when a study was removed, indicating the results of this meta-analysis are robust.

Discussion
In recent years, in spite of improvement in diagnosis and treatment, OC is still characterized by late clinical...
presentation, rapid progression, and poor survival. Considering the high morbidity and mortality of OC, researchers have devoted themselves to identifying new prognostic markers to achieve better clinical decision-making regarding therapy and outcomes. It is well-known that ALDH1 is a CSC marker for a variety of solid tumors, including OC, and is used to enrich ovarian CSCs. ALDH1 is closely related to stemness of ovarian CSCs, including factors such as self-renewal, enhanced invasion, colony formation, and chemoresistance. Some researchers have successfully isolated stem-like cells from OC cell lines by ALDH1 activity and demonstrated their superior abilities in colony formation and tumorigenicity compared to ALDH1 cells in mice. Recently, Kakar et al showed that doxorubicin liposomes, when combined with withaferin A, elicits synergistic effects on inhibition of cell proliferation in OC cells and inhibits

![Figure 5](image-url)

**Figure 5** Sensitivity analysis for all the eligible studies that provided ORs/HRs of high ALDH1 expression for histopathologic grade, FIGO stage, lymph node metastasis, clinicopathological type, DFS, and OS. **Notes:** (A) Histopathologic grade (G1 vs G2+G3). (B) FIGO stage (III–IV vs I–II). (C) Lymph node metastasis (yes vs no). (D) Clinicopathological type (S vs C). (E) DFS (H vs L). (F) OS (H vs L). **Abbreviations:** ALDH1, aldehyde dehydrogenase 1; OR, odds ratio; HR, hazard ratio; FIGO, International Federation of Gynecology and Obstetrics; DFS, disease-free survival; OS, overall survival; S, serous ovarian carcinoma; C, clear cell adenocarcinoma; H, high; L, low.
expression of ALDH1 protein. These observations might indicate that increasing ALDH1 expression is associated with clinicopathological features in patients with OC. Likewise, inhibiting ALDH1 expression could overcome chemoresistance in human OCs. Therefore, ALDH1 could be a potential target for OC therapy.

The association between ALDH1 expression and OC prognosis is being increasingly investigated. Ma and Zhao found that high mRNA expression of five ALDH1 iso-enzymes, including ALDH1A1, ALDH1A2, ALDH1A3, ALDH1B1, and ALDH1L1, was not correlated with OS in serous cancer or endometrioid cancer patients. However, ALDH1A3’s high mRNA expression is associated with worse OS in grade II OC patients. Apart from that, Sun et al also indicated that a high ALDH1 protein expression was associated with poor DFS. Nevertheless, the view from Chang et al was in opposition to this hypothesis. Therefore, we performed a meta-analysis to evaluate the value of ALDH1 as a prognosis marker in OC patients by summarizing all related studies.

The present meta-analysis of 18 included studies, which detected ALDH1 expression in OC tissue samples, indicated that elevated ALDH1 expression was significantly associated with poor OS but not with DFS in patients with ovarian tumor. Additionally, when considering clinicopathological features, our results show that ALDH1 was most frequently elevated in patients with poor clinicopathological characteristics, such as advanced FIGO stage, lymph node metastasis, and distant metastasis, suggesting that ALDH1 may be feasible for tumor staging and metastasis in OC. To sum up, this meta-analysis provides evidence to estimate the significance of ALDH1 detection in patients with OC.

In our meta-analysis, there exist different degrees of heterogeneity for clinicopathological characteristics and survival outcome. Therefore, subgroup analysis was conducted based on sample size and ethnicity. However, subgroup analysis of methods for FIGO stage in ethnicity reached indifferent conclusions. Elevated ALDH1 expression was significantly associated with poor stage in Asians but with early stage in non-Asians. Subgroup analysis of sample size for OS and DFS showed increasing ALDH1 expression in the <100 group rather than the ≥100 group, which might be due to the significant heterogeneity in the ≥100 group. In addition, no evidence indicated publication bias for tumor stage, tumor grade, lymph node metastasis, clinicopathological type, distant metastasis, OS, and DFS in regard to ALDH1 high expression using Begg’s test. In addition, the result of the sensitivity analysis suggested that the pooled OR/HRs were stable and reliable.

There are also several limitations in this current meta-analysis, although our results are promising. First, considerable heterogeneity existed in the present study, to which we applied a relatively conservative random-effects model if there was a significant heterogeneity; therefore, the prognostic value of ALDH1 in OC may be underestimated. Second, due to the current disagreement on the cutoff value for ALDH1, our study did not provide results regarding the optimal cutoff value and whether the cutoff values differed in the assessment of clinicopathological characteristics and prognosis values. In addition, although we performed subgroup analysis based on sample size and ethnicity, other variables were different in the primary studies. The number of included studies was also limited; thus, we could not conduct in-depth subgroup analysis based on the various variables. Therefore, further large-scale multicenter studies on homogeneous patients and diagnostic methods are required to investigate the prognostic values of ALDH1 in OC. Despite the limitations of our study, our meta-analysis is the first study to demonstrate a correlation between ALDH1 and the clinicopathological characteristics and prognosis in OC.

Conclusion
Our results indicate that increasing levels of ALDH1 predict poor prognosis and clinicopathological characteristics in OC. Future studies are needed to explore tailored treatments that directly target ALDH1 for the improvement of survival in OC.

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Author contributions
Changgang Sun was involved in the concept and design of the study. Wenge Zhao drafted the manuscript. All authors participated in acquisition, analysis, and interpretation of the data; critical revision of the manuscript; and read and approved the final version.

Disclosure
The authors report no conflicts of interest in this work.

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