IMF Working Paper

The Costs of Macroprudential Deleveraging in a Liquidity Trap

by Jiaqian Chen, Daria Finocchiaro, Jesper Lindé, and Karl Walentin

IMF Working Papers describe research in progress by the author(s) and are published to elicit comments and to encourage debate. The views expressed in IMF Working Papers are those of the author(s) and do not necessarily represent the views of the IMF, its Executive Board, or IMF management.
IMF Working Paper

Monetary and Capital Markets Department

The Costs of Macroprudential Deleveraging in a Liquidity Trap

Prepared by Jiaqian Chen†, Daria Finocchiaro‡, Jesper Lindé§, and Karl Walentin**

Authorized for distribution by Christopher Erceg

June 2020

Abstract

We examine the effects of various borrower-based macroprudential tools in a New Keynesian environment where both real and nominal interest rates are low. Our model features long-term debt, housing transaction costs and a zero-lower bound constraint on policy rates. We find that the long-term costs, in terms of foregone consumption, of all the macroprudential tools we consider are moderate. Even so, the short-term costs differ dramatically between alternative tools. Specifically, a loan-to-value tightening is more than twice as contractionary compared to loan-to-income tightening when debt is high and monetary policy cannot accommodate.

JEL Classification Numbers:E52, E58

Keywords: Household debt, Zero lower bound, New Keynesian model, Collateral and borrowing constraints, Mortgage interest deductibility, Housing prices.

* The authors are grateful to our discussants Lorenzo Burlon, Stefan Gebauer and Anton Korinek. We are also indebted to Simon Gilchrist and Mathias Trabandt for very helpful feedback. We would like to thank seminar and conference participants at Bank of Israel, Norges Bank, Berlin School of Economics, Federal Reserve Bank of Cleveland, the Fifth NYU Alumni conference, Greater Stockholm Macro Group, Second Annual Workshop of the ESCB Research Cluster on Monetary Economics, Sveriges Riksbank, the Third Research Conference of MMCN, and the ECB’s Fourth Macroprudential Conference. The paper was previously circulated with a different title. The views expressed in this paper are solely the responsibility of the authors and should not be interpreted as rejecting the views of the IMF, Sveriges Riksbank or any other person associated with these organizations. Jacob Ewertzh and Anna Shchekina provided excellent research assistance.

† International Monetary Fund, JChen@IMF.org.
‡ Uppsala University and Sveriges Riksbank, Daria.Finocchiaro@riksbank.se.
§ International Monetary Fund and CEPR, JLinde@IMF.org.
** Sveriges Riksbank, Karl.Walentin@riksbank.se.
1 Introduction

A decade after the unfolding of the Global Financial Crisis (GFC), low nominal and real interest rates have contributed to soaring debt and house prices in many advanced economies like Australia, Canada and several European countries as shown in Figure 1. While record-low interest rates currently imply debt-service-to-income ratios at – or even below – historical levels, policy makers around the world have expressed concerns about households' vulnerability to higher interest rates.¹ Different policies have been put forward to stem the perceived imbalances. While some institutions like the Bank for International Settlements (see e.g. BIS (2016) Box IV.B, pp. 76–77) have advocated leaning against the wind policies on a transient basis, there seems to be consensus that monetary policy should not be the first line of defense to address large persistent upward pressure on housing prices and a desire among households to debt finance housing purchases.²

Amid this background, various macroprudential policies (MPP henceforth), such as caps on loan-to-value (LTV), loan-to-income (LTI) and debt-service-to-income (DSTI) ratios, and fiscal policies, for instance reduced or removed mortgage interest deductibility (MID) and higher property taxes, have been put forth.³ One indication of the activity within borrower-based macroprudential regulation is the fact that 58 different such measures have been implemented in various EU countries in the current decade according to ESRB (2019).

Our aim in this paper is to quantify the short- and long-term output losses associated with different policies that reduce household debt equally in the long term. When quantifying these costs two observations are key. First, as shown in Figure 1, aggregate debt-to-income ratios are highly elevated relative to historical values. Second, in many economies monetary policy is either at, or close to, the effective lower bound without much scope to stimulate the economy following sizeable macroprudential adjustments. These two considerations, which both plausibly are driven by a persistently low equilibrium real interest rate, must be

¹ See Hoffmann et al. (2018).
² See e.g. IMF (2015) and the references therein. Some research have even suggested that leaning against the wind policy may be counterproductive and increase indebtedness in the near-term, see e.g. Gelain et al. (2017).
³ In the following, we call lower mortgage interest deductibility a macroprudential tool because the way we implement the policy avoids any re-distributional effects between borrowers and savers (borrowers fully pay for the mortgage interest rate deductions by lump-sum transfers). Our definition of macroprudential policy is consistent with the one adopted by the ECB, which defines it as any tool which prevents the excessive build-up of risks and smooths the financial cycle over time.
properly taken into account in any rigorous analysis of this issue.

Because current economic conditions are extra-ordinary from a historical perspective, a purely empirical approach to address the question will be of limited value (a large degree of extrapolation would be necessary because few data points cover the situation we are currently facing). Therefore, we use a structural macroeconomic model to perform the analysis. Specifically, we use the workhorse Iacoviello and Neri (2010) two agent New Keynesian model with housing and a collateral constraint, augmented to incorporate long-term debt, housing transaction costs and a broader set of macroprudential instruments. Furthermore, we explicitly take into account the zero (or effective) lower bound constraint on monetary policy. The new elements we introduce have been studied in isolation before, but to the best of our knowledge not jointly yet. Taken together these elements have a considerable impact on the monetary transmission mechanism and allow us to consider an initial position with simultaneously elevated debt levels and monetary handcuffs.

For the countries (except the US) in Figure 1, we observe roughly a doubling of households’ loan-to-income ratios since the 1990s. In our model framework, four factors account for this phenomenon. First, we assume that an increased desire to save has reduced the equilibrium real rate from 3 to 0.5 percent. Such a fall is well in line with the mortgage rate time series in Figure 1 as well as with point estimates from a voluminous empirical literature which aims at measuring equilibrium real interest rates (see e.g., Sajedi and Thwaites (2016), Holston et al. (2017), Lisack et al. (2017), Del Negro et al. (2018) and Fries et al. (2018)). The reduction in the long-term real interest rate contributes to a sizeable decline in the user cost of housing and drives up the real price of houses. The rise in house prices implies a roughly proportional increase in debt and the loan-to-income ratio accordingly rises sharply in our model. Second, we introduce an explicit role for expanded credit supply by loosening the LTV requirement on new loans. This is consistent with the loosening in credit conditions as documented before the GFC. Third, as shown in Figure 1, inflation rates are lower in the 2010’s than in the 1990’s and this further contributes to the increase in LTI in our setting with long-term nominal debt. However, these three factors together cannot fully account for

---

4 Mian and Sufi (2009) reports U.S. microdata evidence in favor of the role of credit supply factors behind the surge in mortgage debt before the crisis. In our experiment, we assume an increase in LTV ratios from 0.75 to 0.85.
the doubling of debt, and we therefore allow for the possibility of a slight increase in home equity extraction to account for the remainder of the increase. Nevertheless, over half of the increase in household indebtedness in our model is accounted for by the decline in real rates.

In our model, low real rates also contribute to a substantial increase in the long-run residential investment as a share of GDP (from 3 to 5 percent of GDP). When the relative price of housing—the user cost—declines, demand for housing increases. Such an increase in demand leads to a surge in residential investment. Interestingly, this increase in residential investment is consistent with the data for several of the countries included in Figure 1. For instance, in Sweden residential investment has risen from less than 2 percent after the housing crisis in the early 1990s to about 7 percent in 2017. Such a structural change and the implied higher exposure to a volatile housing market is a crucial component of the transmission mechanism of macroeconomic disturbances in our framework and poses challenges for stabilization policies. Specifically, in a high debt environment the economy is more volatile not only because the expanded borrowing capacity makes borrowers’ consumption more responsive but also because residential investment constitutes a larger share of output.

Our main findings are as follows. First, the long-term output costs of all the macro-prudential actions in our study are moderate, regardless of whether we consider a steady state with normal (mid-1990s) or elevated debt levels (the current state). Second, the short-term effects of various macroprudential tools depend critically on the initial debt level and the scope for the central bank to provide accommodation. When monetary policy is unconstrained and the initial debt level is low, the short-term output costs will be small (i.e. output does not fall much below its long-term value following an MPP tightening). But in an environment with elevated debt levels and little scope for central banks to cut rates, macroprudential actions may be associated with a significant drop in output and consumption. Specifically, an LTV tightening triggers a large decline in the house price and generates a negative feedback effect on debt capacity and aggregate demand as house prices are part of an LTV constraint. LTI or DSTI tightenings are more efficient tools to curb household

---

5 See Mian and Sufi (2011) for microdata evidence on the importance of home equity extraction for the surge in U.S. mortgage debt.
debt at low output cost. We find that the short-term effects on output and inflation are more than twice as large under an LTV tightening.

To rationalize the use of MPP to reduce indebtedness, we analyze welfare in a nonlinear formulation of our model. Our simulations show that an LTV at 70 percent, i.e. lower than our baseline calibration of 85 percent, is welfare maximizing. The presence of a zero lower bound (ZLB) constraint on the nominal interest rate is a crucial driver behind this finding – disregarding the ZLB constraint the optimal LTV is close to 85 percent. We show that the ZLB constraint is an important reason to restrain household indebtedness as it increases the economy’s vulnerability to disturbances. The welfare analysis illustrates the quantitative significance of the theoretical papers by Lorenzoni (2008), Korinek and Simsek (2016), Davila and Korinek (2018) and Farhi and Werning (2016) which argue that demand and pecuniary externalities associated, respectively, to the zero lower bound and the collateral constraint, provide a rationale to restrain household debt. Taken together, our findings stress that when selecting between various macroprudential tools aimed at stemming household debt, it is important to consider their differing needs for monetary policy accommodation.

Our work is related to different strands of the growing literature using structural macroeconomic models to analyze housing, monetary policy and its interaction with other stabilization policies. Starting with Iacoviello (2005), several papers have explored the linkages between housing, household credit conditions and the macroeconomy (see e.g. Iacoviello and Neri (2010) and Justiniano et al. (2015)). Garriga et al. (2017) show how the presence of multi-period mortgage contracts can enhance the traditional interest-rate transmission mechanism of monetary policy. Differently from ours, their model abstract from nominal rigidities to insulate the effect of the long-term nominal contract aspect of mortgages. Gelain et al. (2017) study the effects of monetary policy in new Keynesian environments with long-term debt. Both of the aforementioned studies abstract from the interaction between monetary and other stabilization policies such as MPP.

A number of empirical papers corroborates some of our findings. Our results mirror the empirical work in Calza et al. (2013), which shows that countries with more developed mortgage markets and high mortgage debt-to-GDP feature larger responses to monetary policy shocks. In their study, the possibility of mortgage equity release and the prevalence
of adjustable rate mortgage (ARM henceforth) contracts turns out to be crucial for the response of consumption. Similarly, using household data for the US and the UK, Cloyne et al. (2018) show that mortgagors’ consumption reacts more strongly than other households’ consumption to monetary policy shocks. Flodén et al. (2018) find that highly indebted Swedish households cut their non-housing expenditures more than less indebted households following changes in the policy rate. In contrast, using state dependent local projections methods Alpanda and Zubairy (2018) show that monetary policy is less effective when the debt is high. They rationalize their findings in a partial equilibrium model where highly indebted households cannot further increase borrowing in response to interest rate cuts. However, taking into account also general equilibrium effects, Walentin (2014) finds opposite results, i.e. the impact of monetary policy is stronger when the level of debt is higher. Richter et al. (2018) provide empirical evidence on the contractionary effects of LTV tightenings, but do not differentiate between situations when monetary policy is constrained and when it is unconstrained (away from the ZLB). They find that a 10 percentage point tightening of the LTV induces a 1.1% reduction in output, although output effects are concentrated in emerging economies. This result is consistent with ours as it is in-between what we find for constrained and unconstrained monetary policy settings. Using a large sample of 63 countries over the period 1991 to 2016, Alam et al. (2019) find that macroprudential tightening through LTV has a significant negative impact on house prices whereas a DSTI/LTI tightening does not have entail any significant effects. This finding provides strong support for our core mechanism that LTV tightenings are more contractionary through their adverse impact on house prices.

Our work also contributes to the growing literature on the interaction between monetary policy and macroprudential regulations (see e.g. Angelini et al. (2014), De Paoli and Paustian (2017), Gelain and Ilbas (2017), Ferrero et al. (2018) and Lambertini et al. (2013)). Like us, Alpanda and Zubairy (2017) build on Iacoviello (2005), but we extend their work by considering the supply side of housing and the ZLB constraint on monetary policy. Their main finding is that monetary policy is too blunt a tool to stabilize households’ debt compared to other more tailor-made housing-related policies. Finocchiaro et al. (2016) and Chen and Columba (2016) both study the long-run effects of deleveraging in models with housing and
a banking sector. On the normative side, both De Paoli and Paustian (2017) and Ferrero et al. (2018) derive a welfare-based loss function in models featuring credit markets frictions. The latter shows that, during boom-bust episodes in housing markets macroprudential policy can help avoid zero lower bound episodes by alleviating debt leveraging. Both Lambertini et al. (2013) and Rubio and Yao (2018) study optimal (countercyclical) LTV rules, for a given average level of LTV. The latter paper shows in a simple stylized model that the macroprudential authority needs to use its instrument more aggressively to stabilize financial cycles when interest rates are low.\(^6\) Differently from these studies, we focus on a positive analysis of the output-debt sacrifice ratio for several macroprudential tools in a quantitatively relevant model with strongly elevated household indebtedness and increased exposure to housing markets. Moreover, in line with empirical evidence on how MPP is conducted (see ESRB (2019)), we consider permanent changes in (borrower-based) MPP tools rather than countercyclical rules.\(^7\)

Our results are consistent with the findings of Mendicino et al. (2018) for bank capital-based macroprudential measures. In their framework, when monetary policy hits the lower bound, it loses the ability of dampening the macroeconomic effects of a (bank) capital requirement increase. Conversely, we focus on borrower-based macroprudential measures and stress the importance of taking into account initial conditions, i.e. the level of debt, to properly assess the trade-offs associated with different tools.

The remaining part of the paper is organized as follows. Section 2 describes our main mechanism in a stylized model. Section 3 presents the quantitative model environment. Section 4 documents the drivers of the increased indebtedness evident in the data. Section 5 then reports the long-term and short-term macroeconomic effects of the various MPP instruments we consider. Section 6 evaluates welfare for different LTV levels accounting for liquidity traps with endogenously determined durations. Finally, Section 7 concludes.

\(^6\) Rubio and Yao (2018) perform their analysis in a model that abstracts from long-term debt, housing investment and business investment. Lambertini et al. (2013) abstract from the ZLB.\(^7\) None of the 58 cases documented in ESRB (2019) is state contingent.
2 Macroprudential deleveraging: inspecting the mechanism

This section illustrates the mechanisms behind our main result, i.e. the stronger near-term contractionary effects under a collateral constraint, in a bare bone version of our model. Our focus here will be on agents with high marginal propensity to consume, i.e. constrained borrowers, whose consumption decisions are crucial for aggregate demand fluctuations. For the ease of exposition, we will abstract from uncertainty, capital accumulation and monetary policy.

Consider a two-period economy where a continuum of identical households can borrow from abroad at a rate $R$, receive a fixed wage every period (normalized to one), and consume housing services ($h$) and non-durable goods ($c$). Agents’ preferences over consumption, housing and hours worked ($n$) are described by

$$
\sum_{t=1}^{2} \beta^{t-1} \left( c_{t} + \log (h_{t}) - \frac{n_{t}^2}{2} \right),
$$

where in line with a borrower-saver framework we assume $\frac{1}{1+R} < \beta < \frac{1}{R}$. Households face the following budget constraint

$$
c_{t} + q_{t} (h_{t} - h_{t-1}) + R b_{t-1} = n_{t} + b_{t},
$$

for given $h_0, b_0 = 0$ and $b_2$.

We consider two set-ups with different specifications for the borrowing limit in the first period. When contracting debt households are either subject to an LTV constraint

$$
b_1 \leq \theta q_1 h_1,
$$

which ties borrowing to the value of housing (see Iacoviello (2005), Farhi and Werning (2016)), or an LTI constraint

$$
b_1 \leq \theta^{LTI} n_1,
$$

which sets the borrowing limit to a fraction of labor income, with $\theta < \beta$. In order to make the two economies comparable in the long-run, we further assume that $\theta^{LTI} = \theta$ and that

---

8 Our parameter restrictions ($\theta < \beta$ and $\beta > 1/(1+R)$) provide sufficient conditions for consumption to be positive and to simply characterize the effects of MPP. Both restrictions are in line with plausible values for the LTV limits, interest rates and discount factors.

©International Monetary Fund. Not for Redistribution
Finally, aggregate housing is in fixed supply, so that in a symmetric equilibrium

\[ h_t = \bar{H}, \forall t. \]

We are now ready to characterize the remaining equilibrium conditions in the LTV and LTI economies. Note that our real rate assumption together with the linearity of utility in consumption cause the borrowing constraints to bind in both environments

\[ \mu = 1 - \beta R > 0, \]

where \( \mu \) is the Kuhn-Tucker multiplier on the borrowing constraint.\(^{10}\)

When the LTV constraint is imposed, house prices, debt and consumption decrease over time. Houses are valued less in the last period since they lose their collateral value. Households’ labor supply is unaffected by the borrowing limit and is constant over time.

When borrowing instead is tied to labor earnings, housing does not have any collateral value and it’s less valued compared to the LTV economy. Under an LTI constraint households instead supply more labor in the first period to be able to borrow more.

### 2.1 Macropreudential tightening

To evaluate the effects of an MPP tightening, we consider a permanent change in \( \theta \)

\[ \theta' = x\theta, \]

where variables with prime refer to the economy which experiences a tightening by setting \( x < 1 \).

By construction, debt will be reduced by the same amount in both economies in the long-run. Even so, the short-run effects of the tightening will differ. In both economies MPP has both a direct and an indirect effect on debt. The direct effect refers to the exogenous change in the regulatory limit \( \theta \). The indirect effect is an amplification channel due to the negative endogenous response of house prices and labor supply in the LTV and LTI economy, respectively. In what follows, we show that movements in house prices are more pronounced

---

\(^9\) This is equivalent to imposing the constraints \( b_{2}^{TV} = \theta q_{2} \bar{H} \) and \( b_{2}^{TI} = \theta \bar{n}_{2} \) where \( \bar{H} \) and \( \bar{n}_{2} \) are aggregate variables since in equilibrium, \( q_{2} = \frac{1}{\bar{m}} \) and \( n_{2} = 1 \), i.e. \( b_{2}^{TV} = b_{2}^{TI} = \theta. \)

\(^{10}\) The complete set of first-order conditions is reported in the Appendix.
than fluctuations in labor supply thereby implying a higher degree of financial amplification in an LTV-economy in the short-term (i.e. period 1).

In the collateral constrained economy, an LTV tightening triggers a contraction in housing demand. Since housing supply is fixed, equilibrium prices will decrease to accommodate the change in demand:

\[ \Delta q_{1,LTV} = -\left[ \frac{x\theta (1 + \beta) \mu}{H (1 - \mu x \theta) (1 - \mu \theta)} \right] < 0, \]

where \( \Delta q \equiv q' - q \), i.e. the difference between the economy experiencing the tightening versus the economy which doesn’t. Households’ borrowing capacity and their consumption will shrink accordingly:

\[ \Delta b_{1,LTV} = -\frac{\theta (1 + \beta) (1 - x)}{(1 - \mu x \theta) (1 - \mu \theta)} < 0, \]
\[ \Delta c_{1,LTV} = \Delta b_{1,LTV}. \]

Interestingly, the drop in prices and consumption is larger the higher the level of initial debt (or equivalently, initial LTV).

Let us now consider the LTI economy. In this case, house prices are unaffected by the policy change

\[ \Delta q_{1,LTI} = 0, \]

while the MPP tightening has an impact on households’ labor decisions, i.e. labor supply in period 1 drops:

\[ \Delta n_{1,LTI} = -\mu \theta (1 - x) < 0. \]

As a result, the total effect on borrowing and consumption is captured by:

\[ \Delta b_{1,LTI} = -\theta (1 - x) (\theta \mu (1 + x) + 1) < 0, \]
\[ \Delta c_{1,LTI} = \Delta n_{1,LTI} + \Delta b_{1,LTI}. \]

To sum-up, the LTI tightening depresses consumption in period 1, since it has a negative impact on both households’ borrowing capacity and their labor supply.

We can now tackle the question at the heart of this experiment, i.e. which of the two MPP tools is most contractionary? It can be easily shown that debt in the first period drops
more after an LTV tightening:

$$\Delta b_{1,\text{LTV}} - \Delta b_{1,\text{LTI}} = -(1 - x) \theta \left[ \frac{\theta^2 \mu^2 ((1 - \theta \mu) x (1 + x) + 1) + \beta}{(1 - \mu x \theta)(1 - \mu \theta)} \right] < 0,$$

as a result of the sharp house price reaction to the MPP change. This translates in a more pronounced fall in consumption:

$$\Delta c_{1,\text{LTV}} - \Delta c_{1,\text{LTI}} = (\Delta b_{\text{LTV}} - \Delta b_{\text{LTI}}) - \Delta n_{1,\text{LTI}}$$

$$= -(1 - x) \theta \left[ \frac{\Xi + \beta - \mu + \theta \mu^2}{(1 - \mu x \theta)(1 - \mu \theta)} \right] < 0$$

where the inequality follows from $$\Xi \equiv \theta^2 \mu^2 ((1 - \theta \mu) x (1 + x + \frac{1}{\theta}) + 1) > 0$$ and that $$\beta > \mu$$ given our restrictions on the parameters.\(^{11}\)

Hence, our simple model demonstrates that deleveraging through LTV tightening has a negative impact on house prices and, via the collateral constraint, on debt and consumption. When debt is instead tied to income via an LTI constraint, a forced deleveraging has less pronounced effects as long as labor supply is relatively more inelastic than house prices to the policy change. In both cases, higher initial debt amplifies the negative responses of consumption to the MPP tightening.

To provide analytical insights, this section has abstracted from nominal rigidities in price and wage formation. However, we will show next that our key finding that an LTV tightening is more contractionary relative to income-based borrowing constraints – and therefore require stronger monetary accommodation – holds up in a full-fledged DSGE model with nominal rigidities.

### 3 The quantitative model

Our quantitative model builds on the Iacoviello and Neri (2010) two-agent new Keynesian economy (TANK) with housing, and is extended to allow for housing transaction costs, long-term household debt, as well as a ZLB constraint for the policy rate. The economy is populated by households and firms. Households consume both housing, $$h$$, and non-housing goods, $$c$$, and provide labor to both sectors. They are divided in two groups with a combined

\(^{11}\)This last condition follows from \(\frac{1}{1 + R} < \beta \Rightarrow 1 - R\beta < \beta.\)
mass of unity; patient households, subscript $P$, and impatient households, subscript $I$, which
discount the future at different rates, $\beta_P > \beta_I$. On the production side, the non-housing
sector combines capital and labor to produce a good that can be used for consumption,
production capital or as an intermediate input in the production of housing. The housing
sector combines capital, labor, land and the intermediate good to produce new housing. For
ease of exposition, in what follows we describe the optimization problems faced by each agent
in the economy and relegate to the Appendix the complete set of first-order conditions and
technical details.

3.1 Households

3.1.1 Patient households

A continuum of identical patient households maximize the following expected lifetime utility
\[
E_0 \sum_{t=0}^{\infty} (\beta_P)^t z_t \left[ \Gamma_c ln(c_{P,t} - \epsilon c_{P,t-1}) + j_P ln(h_{P,t}) - \frac{\omega}{1+\eta} (n_{c,P,t}^{1+\xi} + n_{h,P,t}^{1+\xi}) \right]
\]  
where $c$, $h$, $n_c$, and $n_h$ are consumption, housing, and hours worked in the consumption
and housing sectors, respectively. $z$ and $v$ capture shocks to intertemporal preferences and
labor supply, respectively. $j$ is a housing preference shock aiming to capture preference shifts
towards or away from housing. In the specification of labor disutility, $\eta$ is the inverse Frisch
elasticity of labor supply while $\xi$ is a measure of the labor immobility between sectors, such
that $\xi > 0$ implies that households prefer to spread their working hours to both sectors (see
Hovarth (2000)).

Patient households are the savers in the economy; they accumulate capital and houses
and extend long-term loans to impatient households. Their budget constraint is

\[
c_{P,t} + \frac{i_{c,t}}{A_{k,t}} + i_{h,t} + k_{h,t} + q_t h_{P,t} + p_{il,t} + \frac{L_{P,t}}{P_t} + \frac{a(z_{ct})p_{cn}^{kh}k_{c,t-1}}{A_{k,t}} + a(z_{ht})p_{ct}^{kh}k_{h,t-1} + \phi_{P,t} =
\]

\[
= \frac{w_{c,P,t}n_{c,P,t}}{X_{w,c,t}} + \frac{w_{h,P,t}n_{h,P,t}}{X_{w,h,t}} + (p_{l,t} + R_{l,t})l_{t-1} + \frac{M_{P,t}}{P_t} + R_{c,t}z_{c,t}k_{c,t-1} + R_{h,t}z_{h,t}k_{h,t-1} + p_{b,t}k_{b,t} + q_t(1 - \delta_h)h_{P,t-1} + Div_t
\]

\[12\]

The scaling factor $\Gamma_c = (1 - \epsilon)/(1 - \beta)$ ensures that the marginal utility of consumption is $1/c$ in the
steady state.
where $k_c$, $\delta_{kc}$ and $k_h$, $\delta_{kh}$ denote capital stock and depreciation rates in the non-housing and housing sector, respectively, whereas $k_b$ denotes intermediate inputs and $l$ land.

$A_k$ indicates investment-specific technology in the non-housing sector. $X_{wc}$ and $X_{wh}$ are wage markups accruing to labor unions, $Div$ denotes profits from retail firms and lump-sum payments from labor unions corresponding to the wage markups, $R_c, R_h$ and $R_l$ are rental rates, $z_c$ and $z_h$ capital utilization rates, $a(z_c)$ and $a(z_h)$ utilization costs in terms of capital goods (see the Appendix for functional forms for the utilization costs). The term

$$\phi_{f,t} = \phi_n \left( \frac{h_t}{h_{t-1}} - 1 \right)^2 q_n h_{t-1}, \quad f = \{P, I\}$$

(3)

captures transaction costs borne by households adjusting their housing stock. The law of motion of capital in the two sectors are described by

$$k_{ct} = (1 - \delta_{kc}) k_{ct-1} + F(i_{ct}, i_{ct-1})$$

$$k_{ht} = (1 - \delta_{kh}) k_{ht-1} + F(i_{ht}, i_{ht-1}),$$

where our choice of functional forms for the adjustment costs, $F(\cdot, \cdot)$, is adapted from Christiano et al. (2005), as described in the Appendix.

We follow Alpanda and Zubairy (2017) in the modelling of long-term mortgage contracts. That is, each period savers receive mortgage payments, $\frac{M_t}{P_t}$, which are the sum of interest, $r^M_t$, and constant-principal, $\kappa$, payments$^{13}$,

$$\frac{M_t}{P_t} = [r^M_{t-1} + \kappa] \frac{D_{t-1}}{P_t}.$$  

(4)

The total stock of debt evolves according to

$$\frac{D_t}{P_t} = (1 - \kappa) \frac{D_{t-1}}{P_t} + \frac{L_t}{P_t}.$$  

(5)

It is further assumed that new mortgage loans, $L$, carry a fixed interest rate, $r^F$, for their duration and a fixed fraction of outstanding loans, $\Phi$, are refinanced at this rate. As a result, the effective mortgage rate $r^M_t$ in eq. (4) is a weighted average of present and past rates:

$$r^M_t = (1 - \Phi) \left( 1 - \frac{L_t}{D_t} \right) r^M_{t-1} + \left[ \frac{L_t}{D_t} + \Phi \left( 1 - \frac{L_t}{D_t} \right) \right] r^F_t.$$  

(6)

$^{13}$ See Garriga et al. (2017) and Gelain et al. (2017) for examples of time-varying amortization rates.
3.1.2 Impatient households

Impatient households’ utility functional form is identical to that of patient households, but they discount the future at a higher rate, $\beta_I < \beta_P$, and assign a higher relative utility to housing $j_I > j_P$. Due to their impatience, they accumulate only the required net worth to finance the down payment on their home and borrow the rest from patient households. They do not accumulate any physical capital. Hence, their budget constraint is

$$c_{I,t} + q_t h_{I,t} + \frac{M_{I,t}}{P_t} + \phi_{I,t} = \frac{w_{c,I,t} n_{c,I,t}}{X_{wc,t}} + \frac{w_{h,I,t} n_{h,I,t}}{X_{wh,t}} + q_t (1 - \delta_h) h_{I,t-1} + \frac{L_{I,t}}{P_t} + Div_{I,t} + \tau_t \frac{M}{P_t} D_{I,t-1} - T_t$$

where $\tau$ captures the partial deductibility of interest payments and $T$ are lump-sum taxes.

To analyze the effects of different borrower-based macroprudential measures, we consider alternative specifications of the borrowing constraint. Our baseline specification is that the impatient households’ borrowing is constrained by the following collateral (LTV) constraint

$$\frac{L_t}{P_t} \leq \theta_{LTV} t [h_I - (1 - \delta_h) h_{I,t-1}] + \gamma \left[ q_t (1 - \delta_h) h_{I,t-1} - (1 - \kappa) \frac{D_{I,t-1}}{P_t} \right]$$

which is adapted from Alpanda and Zubairy (2017). Eq. (8) assumes that impatient households capacity to take up new loans is limited to a fraction $\theta_{LTV}$ of their housing investments as collateral. In addition, the households are allowed to extract a constant fraction $\gamma$ of the available home equity every period.

We also consider two other borrowing constraints. First, the following loan-to-income (LTI) constraint

$$\frac{L_t}{P_t} \leq \theta_{LTI} t [w_{c,I,t} n_{c,I,t} + w_{h,I,t} n_{h,I,t}] + \gamma \left[ q_t (1 - \delta_h) h_{I,t-1} - (1 - \kappa) \frac{D_{I,t-1}}{P_t} \right]$$

implying that the value of impatient households’ new loans cannot exceed a certain multiple, $\theta_{LTI}$, of their total income (recall that they do not have any capital revenues) plus home equity extraction. Second, we consider a debt-service-to-income (DSTI) constraint

$$\frac{L_t}{P_t} \leq \theta_{DSTI} t \frac{w_{c,I,t} n_{c,I,t} + w_{h,I,t} n_{h,I,t}}{(1 - \tau_t) r^F + \kappa} + \gamma \left[ q_t (1 - \delta_h) h_{I,t-1} - (1 - \kappa) \frac{D_{I,t-1}}{P_t} \right]$$

which implies that households’ debt service costs on new loans (interest and amortization) cannot exceed a certain multiple of their income.
For small enough shocks around the steady state, all the constraints will hold with equality. Although it would be interesting to analyze the implications of more than one of the constraints above binding simultaneously (following the work by Greenwald (2018), Grodecka (2017) and Justiniano et al. (2019)), we will assume that the borrowing restrictions bind only one at a time to trace out their partial implications for the effects of an MPP tightening on the economy.

3.2 Firms and technology

3.2.1 Wholesale sector

There is a perfectly competitive wholesale sector where capital and labor are inputs in the production of the non-housing good, $Y_t$, while new houses, $IH_t$, are produced with capital, labor, land and an intermediate input. Firms maximize the following profit function:

$$\frac{Y_t}{X_t} + q_t IH_t - \sum_{i=c,h} w_{it} n_{it} + \sum_{i=c,h} w_{it} n_{i,t} + R_{ct} z_{ct} k_{ct-1} + R_{ht} z_{ht} k_{ht-1} + R_{lt} + p_{bt} k_{bt} \right) \tag{11}$$

subject to the production technologies for goods and new houses, respectively:

$$Y_t = \left( A_{ct} \left( n_{c,P,t}^{\alpha} (n_{c,I,t})^{1-\alpha} \right) \right)^{1-\mu_c} (z_{ct} k_{ct-1})^{\mu_c} \tag{12}$$

$$IH_t = \left( A_{ht} \left( n_{h,P,t}^{\alpha} (n_{h,I,t})^{1-\alpha} \right) \right)^{1-\mu_h-\mu_b-\mu_l} (z_{ht} k_{ht-1})^{\mu_h} (k_{bt})^{\mu_b} (k_{lt})^{\mu_l} \tag{13}$$

where $A_{ct}$ and $A_{ht}$ are the productivity levels in the non-housing and housing sectors, respectively and $\alpha$ is the labor income share of patient households. The production specification above implies complementarity across labor skills from the two types of households.

3.2.2 Retailers

To model price stickiness in the non-housing sector, we assume that monopolistically competitive retailers differentiate the homogenous good. These firms buy homogenous goods at the price $P_{tw}$ and sell them at the price $P_t = X_t P_{tw}$, where $X_t$ is the markup. Retailers face Calvo frictions in their price setting, i.e. each quarter they are allowed to choose a new price with a fixed probability $1 - \xi_p$. It is further assumed that the remaining fraction, $\xi_p$, of firms
partially index their prices by a fraction \( \iota_p \) to past inflation. The resulting Phillips curve for net inflation \( \pi = \ln \Pi \) is:

\[
\pi_t - \iota_p \pi_{t-1} = \beta (E_t \pi_{t+1} - \iota_p \pi_t) - \kappa_p \ln \left( \frac{X_t}{X} \right) + \varepsilon_{\pi,t}
\]

(14)

where \( \kappa_p \equiv (1 - \beta \xi_p) (1 - \xi_p) / \xi_p \) and \( \varepsilon_{\pi,t} \) is an i.i.d. markup shock (see e.g. Smets and Wouters (2007)).

Analogously to prices, nominal wages are sticky. The resulting four wage Phillips curves, one for each sector-household pair, are documented in the Appendix.

### 3.3 Monetary, fiscal and macroprudential policy

Monetary policy is constrained by the zero lower bound and the gross nominal interest rate \( R_t \) follows a simple Taylor-type rule:

\[
R_t = \max \left\{ 1, \hat{R}_t \right\}
\]

(15)

with

\[
\hat{R}_t = R_{t-1}^{\rho_R} \left( R \left( \Pi_t \Pi \right) r_x \left( \frac{Y_t}{Y_t^{\text{flex}}} \right) r_Y \left( \frac{Y_t}{Y_{t-1}} \right)^{r_{\Delta Y}} \right)^{(1-r_{\rho R})} \exp (\varepsilon_{r,t})
\]

(16)

where \( \hat{R}_t \) denotes the gross notional (shadow) interest rate, \( \varepsilon_{r,t} \) is an i.i.d. monetary shock, \( Y_t^{\text{flex}} \) measures flex-price-wage output (so that \( \ln \left( Y_t / Y_t^{\text{flex}} \right) \) is the model consistent output gap), and \( \hat{R} \) is the steady state gross nominal interest rate. In the absence of corrective subsidies, the steady state output is inefficiently low in the flex price-wage equilibrium due to the distortions in the product and labor market. Nevertheless, the presence of the output gap in eq. (16) implies that the central bank internalizes policy changes (e.g. macroprudential policy) that have an effect on potential output.

The government balances its budget period-by-period by financing the interest rate deductions with lump-sum taxes paid by impatient households,

\[
T_t = \tau_{r_t} M D_{t-1} P_t.
\]
3.4 Market clearing

Market clearing for non-housing goods implies:

\[ c_{Pt} + c_{It} + \delta_{cl} / A_{kt} + \delta_{ht} + k_{ht} = Y_t. \]

Similarly, the market clearing condition for houses is given by:

\[ h_{Pt} + h_{It} - (1 - \delta_h) (h_{Pt-1} + h_{It-1}) = IH_t. \]

3.5 Calibration

The calibration of the parameters that apply generally, both in the high and the low debt steady state, are documented in Table 1 and motivated below. First, parameters affecting mainly the steady state, e.g. markups, factor shares and depreciation rates, are set to conventional values in the literature. Some steady state parameters are seen as drivers of the increased household debt between the 1990’s and 2010’s and are described in section 4.

Second, parameters that affect only the dynamics are set to the estimated values in Iacoviello and Neri (2010) except investment adjustment cost estimates that we take from Walentin (2014). There are some additional exceptions that we now describe. Motivated by the bulk of the empirical literature (see e.g. Christiano et al. (2005) and Smets and Wouters (2007)), we use higher consumption habit parameters (0.7) than estimated by Iacoviello and Neri (2010) (0.32 and 0.58) to ensure a more conventional monetary policy transmission mechanism. A voluminous literature has also documented lower sensitivity of prices to product and labor market changes observed since the onset of the financial crisis (see e.g. Del Negro et al. (2015) and Lindé et al. (2016)). We therefore consider a higher degree of stickiness in prices compared to Iacoviello and Neri. Since we want to consider a policy rule which recognizes changes to potential output after a persistent MPP tightening, we include the model consistent output gap in the policy rule. We set its parameter to 0.25, consistent with the view that the output gap carries a large weight in a liquidity trap. Such a weight also often approximates optimal policy well in New Keynesian model environments, see e.g.

\[ ^{14}\text{Iacoviello and Neri (2010) did not allow for investment adjustment costs. We decided to include them to ensure a conventional monetary policy transmission mechanism, which is important in our ZLB environment.} \]
the seminal work by Erceg et al. (2000) and the recent paper by Debortoli et al. (2018). We also use a stronger response of the policy rate to inflation than Iacoviello and Neri (2010). The motivation behind the changes to the monetary policy rule is the basic premise that the central bank adopts a more aggressive policy rule in a low interest environment with less policy space so as to avoid persistent deflationary episodes. Finally, we set the share of savers, $\alpha$, to two thirds, 0.67. This value is slightly higher than the prior used by Iacoviello and Neri, but below their posterior median (0.79). Our slightly larger share of borrowers, $(1 - \alpha)$, than estimated by Iacoviello and Neri is inspired by the recent literature on wealthy hand-to-mouth consumers (see e.g. Kaplan et al. (2018)) and our desire to have a total debt to income ratio in line with the data. Our choice of $\alpha$ will still imply that the model underestimates the total debt to income ratio relative to the data (too many savers without any debt relative to constrained borrowers in the model), but the tension is somewhat moderated.\footnote{Specifically, our model implies an aggregate LTI ratio of 83 percent whereas the model which keeps $\alpha$ at 0.79 has an aggregate LTI of 53 percent. The median aggregate LTI of the countries in Figure 1 is close to 200 percent in 2017.}

Third, we set the parameters related to the behavior of borrowers and savers as follows.\footnote{As we have access to granular data from Sweden about the distribution of debt, this part of the calibration focuses on the Swedish case. Furthermore, Sweden appears representative of the countries included in Figure 1.} The speed of amortization parameter is set to yield 46 years of amortization to be in line with Riksbank calculations on the aggregate amortization rate in recent years for Sweden and the average interest rate fixation period is set to 1 year to match the average for Swedish mortgages.\footnote{Source: Sveriges Riksbank’s own calculations using data from the Swedish FSA mortgage survey. This interest rate fixation period is likely a bit shorter than in other countries, and we have therefore checked that our results are robust with respect to this choice (especially the short-term results in Section 5; the long-term effects in Section 4.3 are unaffected by this parameter).} The calibration target for LTI is based on the median loan size-weighted LTI for new mortgage loans obtained by the Swedish FSA, which amounts to 433\% on average for 2015-2017, when the low level of the interest rate had been established.\footnote{Source: Swedish FSA mortgage survey 2015-2017.} The steady state weight on housing in the utility function for patient households, $j_P$, is set jointly with the corresponding weight for impatient households, $j_I$, to yield residential investment of 3\% of GDP (4\% of private sector absorption) and to match LTI of borrowers of 433/2\% in the 1990’s.\footnote{The model has no government consumption, so before presenting any ratio involving GDP we adjust for...}
done analogously; we simply set the LTI (DSTI) ratio parameter, \( \theta_{t}^{LTI} \), \( \theta_{t}^{DSTI} \) to get the same LTI of borrowers as in the LTV constrained economy, while accounting for the change in the real rate and the change in the inflation rate. Our aim with this approach is to keep the LTI, DSTI and LTV constrained economies comparable.

Fourth, let us explain the parameters related to housing transactions and home equity withdrawal. The housing transaction cost parameter \( \phi_{h} \) in eq. (3) is set to match the peak non-housing consumption response to a monetary policy shock for the 1990’s version of our model. The target is taken from Cloyne et al. (2018). Specifically, it is the average of the U.S. and U.K. peak response for non-durable consumption of borrowers, i.e. a 0.33% response to a 25 bp monetary policy shock (their Figures 7 and 8; our Figure 6).

[TABLE 1 HERE]

4 Accounting for the increased household indebtedness

We now describe how our model can account for the large increase in household indebtedness that occurred from the 1990’s to the 2010’s as shown in Figure 1. As noted in the introduction, we assume four key drivers: lower real interest rates, easier credit conditions, lower inflationary pressures, as well as an increase in home equity withdrawal. Table 2 reports the parametrization of these four drivers. As we have access to granular data from Sweden about the credit supply, the calibration focuses on the Swedish case and hence aim to explain a doubling of the aggregate LTI. Nevertheless, the key driving factors (e.g. lower real and nominal rates) are shared with many other advanced economies and we hence believe that the calibration is of general interest and will note below when this might not be the case.

Figure 1 shows how real mortgage rates have fallen in recent decades in many economies. This decline is more moderate for countries with a large share of fixed rate mortgage rates (such as the US and Denmark), which might explain why these two countries are not at an all-time high in terms of household indebtedness. A voluminous literature, see for instance Del Negro et al. (2018), argues that this low real rate represents a persistent new regime. this by multiplying by private sector absorption/GDP=3/4, i.e. accounting for a government consumption of 25% of GDP.
Following this evidence, we assume a real interest rate in the 2010’s 250bp lower than in the 1990’s. In the model, such a fall in the real interest rate accounts for the main chunk of the indebtedness increase in the data. Motivated by the evidence in Figure 2, we additionally assume a higher maximum LTV. More specifically, we let the LTV increase by 10 percentage points, as in the data. Figure 3 shows the LTI and LTV distributions of new loans in Sweden 2015-2017. It indicates extreme bunching at 85% LTV while no analogous bunching is visible in the LTI dimension. We therefore choose the LTV constrained economy as our benchmark. We also note from Figure 1 that inflation rates are lower in the 2010’s than in the 1990’s. We allow for a lower inflation rate in the current decade by calibrating the steady state inflation of the 2010’s to 1.5% instead of the 2.0% in the 1990’s. This further contributes to the increase in DTI in a setting with long-term nominal debt (as lower inflation acts similarly to lower amortization rates). However, these three factors do not cause average indebtedness to rise quite as much as in the data. In line with evidence in Mian and Sufi (2011) we therefore let an increase in home equity withdrawal (HEW) as a fraction of available home equity soak up the residual in the doubling in indebtedness. This implies increasing the HEW share from 1.5% to 2.1% from the 1990’s to the 2010’s. Both these numbers are in the neighborhood of the empirical evidence for the US, 1.7%, reported by Greenspan and Kennedy (2007). Home equity withdrawal appear similar between U.S. and Sweden in terms of frequency and amounts according to Li and Zhang (2017). Jointly, these four changes double the LTI of borrowers, from 217% to 433%.

The decomposition between the debt drivers are as follows. Changing only the interest rate accounts for 54% of the increase in borrowers DTI. Changing also the LTV ratio contributes an additional 24% of the increase and the change in the HEW accounts for 12% of the increase. Finally, the lower inflation accounts for 10% of the increase in borrowers’ DTI.

| TABLE 2 HERE |

Table 3 compares the long-run equilibrium in the 1990’s and the 2010’s. The table shows that DSTI ratios of borrowers increase substantially less than the increase in debt (and LTI) while the pure interest payment over income (not including amortization) actually decrease slightly. The latter follows from that households wish to spend a fixed fraction of
their expenses on housing due to Cobb-Douglas preferences over housing and non-housing consumption combined with the fact that the user cost of housing consists of both the real (after tax) interest rate and housing depreciation. We also note that the lower real rate increases the share of GDP accounted for by both non-residential and residential investment. The increase in the latter is most consequential as the share almost doubles between the 1990’s and the 2010’s. As a result, the economy with a lower real rate is much more exposed to fluctuations in the housing market.

Third, the doubling in the LTI ratio between 1990’s and 2010’s imply an increase in real house prices of roughly one third. This is less than in the data, but partially reflects the higher residential investment and the associated sizeable increase in the housing stock in the model. Hence, the value of the housing stock relative to GDP rises notably more, by 110 percent, closer in line with the real house price increase observed in the data (slightly more than a 100 percent increase according to Figure 1).

Fourth, a final take-away from Table 3 is that the steady state properties of the model are approximately invariant to the type of borrowing constraint.\(^{20}\)

\(\text{TABLE 3 HERE}\)

5 Effects of macroprudential tightening

In this section we characterize the effects of four borrower-based policies aimed at reducing the aggregate loan-to-income ratio by 10.2%: a tightening of the LTV, LTI and DSTI constraints, and a reduction in MID from a starting point of 30% (given an LTV constrained economy).

5.1 Long-term effects

Table 4 reports the long-term effects of tightening the various macroprudential instruments. The reductions in the LTV, LTI and DSTI parameters are set to imply an identical decline in

\(^{20}\) Regarding the dynamics, Figure A.1 in the Appendix documents the monetary policy shock impulse responses across models (i.e., for different borrowing constraints) in a high debt environment. The main take-away from this figure is that the monetary transmission mechanism is quite similar across models, with the exception that in the LTV model borrowers’ consumption contracts slightly more.
the aggregate long-run loan-to-income ratio (10.2%) as obtained when removing MID in the high-debt LTV-constrained economy. This scaling is done separately for the low indebtedness and the high indebtedness regime. In the low indebtedness regime the LTV ratio is reduced from 0.75 to 0.65 and in the high indebtedness regime it is reduced from 0.85 to 0.77.\textsuperscript{21}

[TABLE 4 HERE]

The aggregate long-term effects of all the four policies on output and consumption are moderate. Output falls in response to all debt-reducing policies under consideration, even if the mechanisms are different. When a collateral constraint is binding (LTV and MID tightening), reducing debt capacity has a small, positive effect on borrowers consumption, while housing investment and house prices instead strongly contract. The large housing decrease of borrowers in response to a tighter LTV constraint stems mainly from the decreased usefulness of housing as collateral as discussed previously in the stylized model (Section 2). When MID is removed, debt-financed housing effectively (after tax) becomes more expensive and borrowers accordingly reduce their housing stock. In the long run, borrowers instead devote a higher share of their income to non-housing consumption. The resulting fall in their marginal value of wealth induces borrowers to work less, thereby causing the observed contraction in output.

On the contrary, when debt is constrained by labor income (LTI and DSTI tightening), the fall in house prices and housing investment is much smaller as explained in our analysis with the stylized model in Section 2. In the high-debt economy, borrowers will also reduce their non-housing consumption in response to an LTI tightening, and their labor supply decreases as labor income loses part of its “collateral” value. Specifically, when households are allowed to borrow in proportion to their labor income, the borrowing limit has a direct, positive, impact on their labor supply as impatient households work more to be able to borrow more.

Comparing tightening of LTI and DSTI constraints, we notice that they have basically identical effects on all variables in the long run. The result reflects that the only difference between these two constraints is that DSTI takes into account the time-variation in the

\textsuperscript{21} In the low debt environment, mortgage deductibility is only reduced to 6.35\% to obtain a 10.2\% reduction in the aggregate loan-to-income ratio.
nominal interest rate (and neither an LTI nor DSTI change will affect the nominal rate in the long run). Finally, we note that the long-term aggregate effects on the economy of the debt reducing policies are roughly invariant, in percentage terms, to the debt level. So the basic mechanisms at work in the long-run are not contingent on the initial debt level.

5.2 Short-run effects

We now turn to analyzing the short-run effects of the various MPP tools. First, we focus on the macroeconomic effects of an LTV tightening in a low vs. high debt economy when the economy is in a liquidity trap. Next, we compare the effects of an LTV cut with those obtained for alternative borrower-based MPP tools (LTI/DSTI constraints and MID). However, we begin with explaining how we generated a baseline in which the economy is driven to the zero lower bound (ZLB).

5.2.1 Simulations set-up

The economy is assumed to be driven to the ZLB by a mix of adverse shocks which lowers inflation below the central bank’s targeted rate and drives down output below its potential level. In our setting in which the only nonlinearity is the ZLB constraint, there is no need to specify which particular shock(s) have driven the economy to the ZLB. The only thing that matters is the path of the notional, or shadow rate $\hat{R}_t$ in eq. (16), as the path of this variable determines the expected duration of the ZLB which we assume is two years absent any MPP changes.\footnote{For proof, see e.g. Erceg and Lindé (2014).} When we add the various MPP tools to the baseline scenario, the duration of the liquidity trap is kept unchanged. This means that we consider the marginal effects (ZLB duration fixed) as opposed to the average effect (ZLB duration extended to the extent that the MPP impacts the economy adversely). Hence, the implicit assumption is that the central bank is able to cut rates as much as needed after the two years it is constrained by the ZLB in the baseline scenario without any MPP measures.\footnote{See Erceg and Lindé (2014) for a detailed discussion of the difference between marginal and average impulses in a liquidity trap.} Had we instead focused on the average effects and allowed the duration of the liquidity trap to change in response to
the MPP actions, our results would have further strengthened as the difference between LTV and LTI/MID would be even greater. In light of this, the results we report are conservative.

We implement the short-run experiments for debt reduction in the following way. The change in the LTV, LTI, and DSTI parameters is modelled as an AR(1) with a coefficient arbitrarily close to unity. Hence, we think about these changes as structural and not driven by any rationale to fine-tune cyclical variation in indebtedness or the economy. Noting that the LTV/LTI/DSTI constraint only applies to new mortgage loans while mortgage deductibility pertains to the entire stock of outstanding loans, we use an AR(2)-process for the latter to obtain a gradual reduction of MID \( (\tau_t) \) and thereby a similar path for household debt as an LTV tightening. In particular, \( \tau_t \) in eq. (7) follows

\[
\Delta \tau_t = \rho_{r,1} \Delta \tau_{t-1} - \rho_{r,2} (\tau_{t-1} - \tau) + \varepsilon_{r,t},
\]

where we set \( \rho_{r,1} = 0.9 \) and \( \rho_{r,2} = 0.0000001 \) to achieve a near-permanent gradual decline in \( \tau \).

### 5.2.2 Tightening of LTV

Panel A in Figure 4 reports the aggregate responses to the reduction in the LTV ratio when monetary policy is constrained by the ZLB for 8 quarters, for a high and low debt economy, respectively. The message of this panel is twofold: i) The economy’s transition in response to an LTV tightening implies large costs in terms of inflation and output when monetary policy is constrained, and ii) these costs are roughly three times as large in the high debt economy. In particular, GDP (inflation) will initially fall by more than 3 (less than 1) percent in the high debt economy in response to an LTV tightening that reduces LTI by 10.2% in the long run.\(^\text{24}\) The corresponding numbers for the low debt economy is a 0.8 percent reduction in GDP and a 0.3 percent reduction in inflation.

To understand the drivers behind this large contraction in the high debt economy, Figure 5 reports the disaggregate effects of the LTV cut. The LTV tightening has a direct effect on borrowers’ borrowing capacity and therefore their demand. There is also an indirect strong endogenous contractionary effect on their debt capacity as house prices fall by more than \(\text{24 That is, a reduction in borrowers’ LTI from 433 to about 390 percentage points in the high debt economy.}\)
5 percent on impact (vs. 2.6 percent in the low debt economy). In addition, there is a further endogenous reduction of borrowing capacity as borrowers reduce their housing stock by roughly 1 percent on impact (and then gradually more over time). We note that all three major components of aggregate demand fall in response to the LTV reduction. Residential investment is most adversely affected, followed by aggregate (non-housing) consumption and non-residential investment.\(^{25}\) The fall in aggregate consumption is due to the large initial contraction in borrowers’ consumption. We note that this reduction in borrowers’ consumption is four times as large in the high debt regime whereas the reduction in borrowers’ housing stock is similar across debt regimes. Quantitatively, the housing transaction costs are important for this response as they influence borrowers’ willingness to off-load housing to savers in the short run. An additional reason for the gradual decline of the borrowers’ debt and housing stock is that the tightened LTV only applies to new loans.

Two main forces are behind our finding that an LTV tightening contracts GDP three times more in the high debt economy. First, the LTV tightening requires more monetary accommodation in a high debt economy. To show this, Panel B in Figure 4 reports the aggregate responses to the identical LTV cut when monetary policy is unconstrained. Note that the unconstrained policy rate response is three times larger compared to the low debt economy.\(^{26}\) Second, monetary policy is more potent in a high debt environment. This has the unpleasant implication that a ZLB is a particularly detrimental constraint in such an environment. To show this, Figure 6 reports the aggregate effects of a hike in the policy rate with 25 basis points for both high and low debt economies. The figure documents that the inflation response to a monetary shock is only mildly amplified in the high debt regime. Even so, output responds much stronger (by roughly 50%) to monetary policy shocks in the high debt regime. The reason that output responds stronger in that regime is the stronger contribution from residential investment as well as borrowers’ consumption. Figure 6 report the main demand components in terms of their contribution to GDP to make responses comparable between indebtedness regimes; as documented in Table 3 both forms

---

\(^{25}\) The strength in the decrease in residential investment might appear extreme. But properly accounting for the volatility of this variable moderates this impression. There are many occasions in the recent history when residential investment has fallen by more than 10 percent within a year, see Figure 1.\(^{26}\) We are not describing optimal monetary policy, but simply noting that the Taylor rule used seem appropriate in the sense that it trades off deviations in inflation against deviations in output in response to the LTV tightening.
of investment constitute a substantially larger share of GDP in the high debt regime.

5.2.3 Short run effects of income-based MPP tools

We now relate our previous findings to income-based MPP tools and show that the impact of the latter can differ substantially in the short run. Specifically, when the economy is highly exposed to housing markets and monetary policy cannot accommodate, the short-term macroeconomic costs of an LTI or an DSTI tightening are substantially lower than from an LTV tightening.

Figure 7 reports the aggregate responses to various MPP tools – sized to achieve the same long-term LTI reduction – when monetary policy is constrained by the ZLB for two years and indebtedness is high. As can be seen in Figure 7, the output response is more than twice as strong in the case of tightening the LTV constraint. The reason for the stronger output contraction yielded by a tightening of the LTV constraint relative to LTI/DSTI is mainly the negative feedback effects that occur through house prices in the presence of an LTV constraint. We demonstrated this qualitatively already in our simple two period model. Figure 8 quantifies these dynamics. Differently from the LTI and DSTI constraints, a tightening in LTV constrained economy triggers a large drop in house prices which reduce the borrowing capacity of households. This reduced borrowing capacity feeds back to borrowers’ demand for both housing and non-housing consumption. The result of this is that an LTV tightening results in substantial “over-shooting” in terms of the reduction in all demand components for the first two years of tightened LTV. The corresponding negative feedback through decreased income in the case of an LTI or DSTI constraint is much weaker as income is less volatile than house prices. One reason for the muted response in borrowers’ labor income in the LTI/DSTI case is that borrowers optimally are able to increase their short-term labor supply in response to an LTI/DSTI tightening so as to smooth their consumption.

The amplification mechanism in our LTV model is not new in the literature. It is simply another incarnation of the financial accelerator outlined by Bernanke et al. (1998), Kiyotaki and Moore (1997) and Iacoviello (2005).

27 The figure also includes MID removal which will be discussed in section 5.2.4.
5.2.4 Short run effects of an MID removal

Finally, we show that our model implies that even gradual MID removal can entail significant macroeconomic costs when interest rates are low and LTV constraints are binding.

The green line in Figures 7 and 8 report the results of a gradual MID removal when monetary policy is constrained. This puts downward pressure on house prices as it increases the effective (after tax) user cost of debt-financed housing. Even though MID is only gradually removed, the house price fall occurs immediately due to the forward-looking nature of asset prices. The reason that the macroeconomy contracts strongly is that the fall in house prices reduces borrowing capacity because of the LTV constraint. Quantitatively the macroeconomic contraction is almost as strong as for an LTV tightening. On the contrary, a removal of MID under LTI or DSTI constraints would require no monetary accommodation and only gradually reduce output as shown in Figure A.2 in the Appendix.

The main take-away from this analysis is that removing mortgage deductibility under LTV constraints can be associated with large short run macroeconomic costs even in a environment with low current interest rates if households are forward-looking and expects interest rates to rise within a foreseeable future. So even in this case, more favorable short-term effects are obtained by imposing LTI or DSTI constraints when deleveraging.

6 Welfare implications of reducing debt limits

A possible critique of our analysis so far is the absence of a welfare justification for using macroprudential tools, e.g. an LTV tightening, to reduce household indebtedness. In fact, this type of critique is often voiced in response to proposals for increased macroprudential regulation. In this section we tackle this issue and assess the welfare gains of tightening the LTV constraint, eq. (8) both for society as a whole and separately for the two types of households. We perform this exercise in our high debt (low steady state real and nominal rate) economy. To perform this analysis, we follow Schmitt-Grohe and Uribe (2007) and approximate lifetime utility up to second-order. Specifically, we compute a second-order approximation of the model around its non-stochastic steady state and simulate the pruned system (following Andreasen et al. (2018)) for a large number of time periods (5,000) for
different LTV values. Importantly, we use exactly the same innovations to all estimated shock process for each LTV value we consider. This way, we parse out the partial derivative of the steady state LTV value in the borrowing constraint.

We then consider the unconditional welfare ($W_P$ and $W_I$) of the two groups of households in our economy by computing the stochastic mean of their discounted sum of utility along the equilibrium path. Following Lambertini et al. (2013), we also construct a measure of social welfare

$$W_S = (1 - \beta_P) W_P + (1 - \beta_I) W_I$$

where the chosen social weights assign equal welfare to the two groups of households for the same constant consumption stream.

Two alternative assumptions about the conduct of monetary policy are considered. As a benchmark, we assume that the central bank is unconstrained when setting policy rates. Under this assumption, the central bank can always set the policy rate as low as it deems appropriate according to the interest rate rule eq. (16), and abstract from the ZLB in eq. (15). In our second case, we assume that the central bank is constrained by an effective lower bound (zero here) on policy rates when setting the policy rate. To impose the ZLB on the second-order system, we follow Hebden et al. (2010) and add positive current and anticipated monetary policy shocks when the ZLB binds. Under this procedure, the current and expected policy rate is non-negative in each state.

The presence of a ZLB constraint amplifies the welfare costs related to high indebtedness. A higher LTV will trigger an increase in macro volatility and hence in the frequency of hitting the ZLB constraint. This last channel entails larger welfare costs and in low interest rate environment these costs are substantial.

The estimated shock processes are all adapted from Iacoviello and Neri (2010), with one exception. Since we are considering a low-interest environment (i.e. a steady state nominal interest rate of 2 percent instead of 5 percent as in Iacoviello and Neri (2010)), we allow for the possibility that the variance of the estimated monetary policy shock is lower than its historical average. Specifically, we shrink the standard deviation of the monetary policy shock by the ratio of the steady state nominal interest rate in our model and the average Federal funds rate in Iacoviello and Neri (2010)’s sample (1965-2006). In doing so we obtain
a modified standard deviation of monetary policy shocks of about 34 percent of the value estimated in that paper. Had we used the historical estimated value, the welfare costs when imposing the ZLB on the policy rate would have been significantly larger and the optimal unconditional LTV value notably lower.

The results of the welfare calculations are reported in Figure 9. The upper panels show (from left to right) total, savers and borrower average welfare as function of the steady state LTV level, \( \theta^{LTV} \), in eq. (8). As noted previously, we present results when the ZLB is imposed (red line with *'s) and when the ZLB is not imposed (blue line with x’ s).

When monetary policy in unconstrained, the unconditional social welfare is maximized for an LTV of about 0.85, which is the benchmark value in our high indebtedness model (to be exact, maximum welfare is obtained for an LTV of 0.84). From the second panel we see that the savers welfare is monotonically increasing in the LTV value, whereas the third panel shows that borrower’s welfare is deteriorating when the steady state LTV exceeds 0.77. Hence, the LTV level which maximizes total welfare trades off the utility gains the savers obtain with a higher LTV with the utility costs the borrowers experience for higher LTV values. Higher LTVs have a negative impact on borrowers’ welfare as it causes an increase in their consumption volatility. For relatively low values of the LTV, the welfare of the savers dominates but for higher LTVs the increasing rate of deterioration of borrowers’ welfare dominates.

When the ZLB is imposed, however, total welfare is maximized for a notably lower LTV value (0.71) in our model. As seen from the second and third panel, the lower optimal LTV value is entirely driven by lower welfare of borrowers for higher LTV values. Hence, a tightening of macroprudential policy from an LTV value of 0.85 to 0.75 can be rationalized on unconditional welfare grounds in an environment where the ZLB is taken into account.

Our results indicate that borrowers are more negatively affected by the ZLB than savers. To shed light on this issue, the lower panels in Figure 9 show the simulated distributions for the output gap (left), non-housing consumption of savers (middle) and borrowers (right) for an LTV fixed at 0.70. As can be seen here, the ZLB introduces a severe downward skew for the output gap and non-housing consumption of borrowers, which is very costly in terms of welfare. Savers non-housing consumption, on the other hand, is largely unaffected by
the ZLB, presumably because the savers can use their savings to smooth their non-housing consumption and gain from buying cheaper houses from the borrowers when the ZLB binds.

Finally, we note that our analysis implicitly assumes that neither unconventional monetary policy (UMP) nor fiscal policy tools are available to stabilize the economy when the ZLB binds. Given the existence of UMPs, this assumption may bias the results in favor of a too low optimal LTV value. On the other hand, there is a considerable debate on the effectiveness of UMP to deal with deep recessions (see e.g. CGFS Report (2019) and the references therein) and whether fiscal policy can act quickly and forcefully enough to provide meaningful relief in recessions. Moreover, it should be noted that the ZLB does not bind unreasonably often in our simulations: the economy is only at the ZLB in about 300 out of 5,000 simulated periods when LTV equals 0.7 (i.e. with a probability of 0.06). Hence, our finding of a lower optimal LTV ratio when policy rates are subject to an effective lower bound may hold up well even if we allow for fiscal policy and UMP unless those tools can be designed to put money directly in the hands of constrained borrowers.

[Figure 9 HERE]

7 Conclusions

In this paper we have documented four factors that together can account for a doubling of household indebtedness between the 1990’s and the 2010’s in several advanced economies, with the lower real mortgage rate as the main driver.

We have shown that in the presence of an effective lower bound on policy rates, there is a need to limit household indebtedness from a welfare perspective even if house prices are consistent with fundamental values. Specifically, when the policy rate is close to the ZLB, our model implies that unconditional welfare is maximized when LTV is capped around 70 percent. The limited ability of the central bank to provide monetary stimulus in a recession is an important reason to restrict household borrowing as monetary handcuffs imply higher volatility of borrowers’ non-housing and housing consumption in settings with high LTV.

28 see Biljanovska et al. (2019) for a model considering optimal macroprudential policy in the presence of asset price bubbles.
But how should household indebtedness be reduced when debt is already above its welfare maximizing level to begin with? To study this issue, we compared both short- and long-term effects of a number of borrower-based macroprudential tools. We find that the long-term output costs of all the macroprudential measures (tightening of LTV, LTI or DSTI ratios or lower mortgage interest deductibility) are moderate. The short-term effects, however, depend critically on which MPP tool that is used. In an environment with elevated debt levels and little scope for the central bank to cut rates, an LTV tightening or reduction in mortgage deductibility may be associated with a significant drop in output and consumption in the near-term as they generate a large housing price decline which reinforces a transitory negative feedback effect on borrowing capacity as prices are part of the collateral constraint. Under these initial conditions, LTI or DSTI tightenings are more efficient tools to curb household debt at lower output cost as these tools avoid the adverse feedback effects through the collateral constraint and hence are less dependent on monetary accommodation. When the initial debt level is lower and monetary policy is unconstrained, the short-term costs will be smaller and any of the four MPP tools studied in this paper can be used to reduce indebtedness at low output cost.

We leave some important extensions for future research. For example, although our theoretical framework is quite rich, it does not allow for heterogeneity between borrowers. It would be interesting to examine the robustness of our policy conclusions in a framework that does allow for such heterogeneity. It is conceivable that a such a framework would imply that a combination of tools, with different tools binding for different borrowers, may entail lower short-term costs or better address long-term financial stability risks. Allowing for borrower heterogeneity may also show that macroprudential deleveraging engineered solely through lower loan-to-income limits may require shutting down access to credit disproportionally for low income borrowers.

All told, our findings stress that when designing macroprudential policies aimed at addressing current household debt imbalances, it is critical to account for their interaction with monetary policy. Extending the model to heterogeneous borrowers with different binding borrowing constraints would improve the welfare analysis although it most likely would not change the main results in the paper.
References

Alam, Zohair, Adrian Alter, Jesse Eiseman, Gaston Gelos, Heedon Kang, Machiko Narita, Erlend Nier, Naixi Wang, 2019, “Digging Deeper–Evidence on the Effects of Macroprudential Policies from a New Database,” IMF Working Papers WP/19/66.

Alpanda, Sami and Sarah Zubairy, 2017, “Addressing household indebtedness: Monetary, fiscal or macroprudential policy?,” European Economic Review, 92(C), 47–73.

Alpanda, Sami and Sarah Zubairy, 2019, “Household Debt Overhang and Transmission of Monetary Policy,” Journal of Money, Credit and Banking, 51(5), 1265–1307.

Andreasen, Martin, Jesús Fernández-Villaverde and Juan F. Rubio-Ramirez (2018), “The Pruned State-Space System for Non-Linear DSGE Models: Theory and Empirical Applications,” Review of Economic Studies 85, 1–49.

Angelini, Paolo, Stefano Neri and Fabio Panetta, 2014, “The Interaction between Capital Requirements and Monetary Policy,” Journal of Money, Credit and Banking, 46(6), 1073–1112.

Bernanke, Ben, Mark Gertler and Simon Gilchrist, 1998, “The Financial Accelerator in a Quantitative Business Cycle Framework,” in Handbook of Macroeconomics, Volume 1C, Handbooks in Economics, Amsterdam: Elsevier, 1341–93.

Nina Biljanovska, Lucyna Gornicka and Alexandros Vardoulakis, 2019, “Optimal Macroprudential Policy and Asset Price Bubbles,” IMF Working Papers WP/19/184.

Bank for International Settlements, 2016, 86th Annual Report, Bank for International Settlements, www.bis.org.

Calza, Alessandro, Tommaso Monacelli and Livio Stracca, 2013, “Housing Finance and Monetary Policy,” Journal of the European Economic Association, 11, 101–122.

CGFS Report, 2019, “Unconventional Monetary Policy Tools: A Cross-Country Analysis,” Committee on the Global Financial System Papers No. 63, Bank for International Settlements.
Chen, Jiaqian, and Francesco Columba, 2016, “Macroprudential and Monetary Policy Interactions in a DSGE Model for Sweden,” IMF Working Paper 16/74.

Christiano, Lawrence J., Martin Eichenbaum and Charles L. Evans, 2005, “Nominal Rigidities and the Dynamic Effects of a Shock to Monetary Policy,” *Journal of Political Economy*, 113(1), 1–45.

Cloyne, James, Clodomiro Ferreira and Paolo Surico, 2020, “Monetary Policy when Households have Debt: New Evidence on the Transmission Mechanism”, *Review of Economic Studies*, 87(1), 102–129.

Dávila, Eduardo and Anton Korinek, 2018, “Pecuniary Externalities in Economies with Financial Frictions”, *Review of Economic Studies* 85(1), 352–395.

Debortoli, Davide, Jinill Kim, Jesper Lindé and Ricardo Nunes, 2019, “Designing a Simple Loss Function for Central Banks: Does a Dual Mandate Make Sense?,” *Economic Journal*, 129(621), 2010–2038.

Del Negro, Marco, Marc P. Giannoni and Frank Schorfheide, 2015, “Inflation in the Great Recession and New Keynesian Models.” *American Economic Journal: Macroeconomics*, 7(1), 168–196.

Del Negro, Marco, Domenico Giannone, Marc P. Giannoni and Andrea Tambalotti, 2019, “Global Trends in Interest Rates”, *Journal of International Economics*, Elsevier, 118(C), 248–262.

De Paoli, Bianca and Mathias Paustian, 2017, “Coordinating Monetary and Macroprudential Policies,” *Journal of Money, Credit and Banking*, 49, 319–349.

Eggertsson, Gauti and Paul Krugman, 2012, “Debt, deleveraging, and the liquidity trap: a Fisher–Minsky–Koo approach,” *Quarterly Journal of Economics*, 127(3), 1469–1513.

Erceg, Christopher J., Dale W. Henderson and Andrew T. Levin, 2000, “Optimal Monetary Policy with Staggered Wage and Price Contracts.” *Journal of Monetary Economics*, 46, 281–313.
Erceg, Christopher J. and Jesper Lindé, 2014, “Is there a fiscal free lunch in a liquidity trap?,” *Journal of the European Economic Association*, 12, 73–107.

European Systemic Risk Board, 2019, “National measures of macroprudential interest in the EU/EEA”, mimeo.

Farhi, Emmanuel and Iván Werning, 2016, “A Theory of Macroprudential Policies in the Presence of Nominal Rigidities,” *Econometrica*, 84, 1645–1704.

Ferrero, Andrea, Richard Harrison and Benjamin Nelson, 2018, “Concerted Efforts? Monetary and Macro-Prudential Policies,” Bank of England Staff Working Paper No. 727.

Finocchiaro, Daria, Magnus Jonsson, Christian Nilsson and Ingvar Strid, 2016, “Macroeconomic effects of reducing household debt”, *Sveriges Riksbank Economic Review* 2016:02.

Flodén, Martin, Matilda Kilström, Jósef Sigurdsson, and Roine Vestman, 2018, “Household Debt and Monetary Policy: Revealing the Cash-Flow Channel,” mimeo, Stockholm University.

Fries, Sébastien, Jean-Stéphane Mésonnier, Sarah Mouabbi, and Jean-Paul Renne, 2018, “National natural rates of interest and the single monetary policy in the euro area,” *Journal of Applied Econometrics*, 33(6), 763–779.

Garriga, Carlos, Finn E. Kydland, and Roman Šustek, 2017, “Mortgages and Monetary Policy”, *Review of Financial Studies*, 30(10), 3337–3375.

Gelain, Paolo, Kevin J. Lansing and Gisle James Natvik, 2018, “Leaning Against the Credit Cycle”, *Journal of the European Economic Association*, 16(5), 1350–1393.

Gelain, Paolo and Pelin Ilbas, 2017, “Monetary and macroprudential policies in an estimated model with financial intermediation,” *Journal of Economic Dynamics and Control*, 78, 164–189

Greenspan, Alan and James Kennedy, 2007, “Sources and Uses of Equity Extracted from Homes”, *Oxford Review of Economic Policy*, 24(1), 120–144.
Greenwald, Daniel, 2018, “The Mortgage Credit Channel of Macroeconomic Transmission”, mimeo.

Grodecka, Anna, 2019, “On the effectiveness of loan to-value regulation in a multiconstraint framework”, *Journal of Money, Credit and Banking*, forthcoming.

Hebden, James S., Jesper Lindé, and Lars E.O. Svensson (2010), “Optimal Monetary Policy in The Hybrid New-Keynesian Model Under The Zero Lower Bound,” *Manuscript*, Federal Reserve Board.

Hoffmann, Peter, Sam Langfield, Federico Pierobon and Guillaume Vuillemey, 2018, “Who Bears Interest Rate Risk?,” *Review of Financial Studies*, 32(8), 2921–2954.

Holston, Kathryn, Thomas Laubach and John C. Williams, 2017, “Measuring the natural rate of interest: International trends and determinants,” *Journal of International Economics* 108, 59–75.

Horvath, Michael, 2000, “Sectoral Shocks and Aggregate Fluctuations,” *Journal of Monetary Economics*, 45(1), 69–106.

Iacoviello, Matteo, 2005, “House Prices, Borrowing Constraints, and Monetary Policy in the Business Cycle,” *American Economic Review*, 95(3), 739–764.

Iacoviello, Matteo and Stefano Neri, 2010, “Housing Market Spillovers: Evidence from an Estimated DSGE Model,” *American Economic Journal: Macroeconomics*, 2(2), 125–64.

IMF 2015, “Monetary Policy and Financial Stability,” IMF Staff Report, August.

Justiniano, Alejandro, Giorgio Primiceri and Andrea Tambalotti, 2015, “Household leveraging and deleveraging,” *Review of Economic Dynamics*, 18(1), 3–20.

Justiniano, Alejandro Giorgio Primiceri and Andrea Tambalotti, 2019, “Credit Supply and the Housing Boom”, *Journal of Political Economy*, 127(3), 1317–1350.

Kaplan, Greg, Benjamin Moll and Giovanni L. Violante, 2018, “Monetary Policy According to HANK”, *American Economic Review*, 108(3), 697–743.
Kiyotaki, Nobuhiro and John Moore, 1997, “Credit Cycles”, *Journal of Political Economy*, 105(2), 211–248.

Korinek, Anton and Alp Simsek, 2016, “Liquidity Trap and Excessive Leverage”, *American Economic Review* 106(3), 699–738.

Lambertini, Luisa and Maria Teresa Punzi, 2013, “Leaning Against Boom-Bust Cycles in Credit and Housing Prices: Monetary and Macroprudential Policy,” *Journal of Economic Dynamics and Control*, 37(8), 1500–1522.

Li, Jieying and Xin Zhang, 2017, “House Prices, Home Equity, and Personal Debt Composition”, Sveriges Riksbank Working Paper Series No. 343.

Lindé, Jesper, Frank Smets and Rafael Wouters, 2016, “Challenges for Central Banks’ Macro Models”, Chapter 28 in John B. Taylor and Harald Uhlig (Eds.), Handbook of Macroeconomics Vol. 2, North-Holland Elsevier Science, New York.

Lisack, Noemie, Rana Sajedi and Gregory Thwaites, 2017, “Demographic trends and the real interest rate,” Bank of England Staff Working Paper no. 701.

Guido Lorenzoni, 2008, “Inefficient Credit Booms,” *Review of Economic Studies*, 75(3), 809–833.

Mendicino, Caterina, Kalin Nikolov, Javier Suarez and Dominik Supera, 2019, “Bank Capital in the Short and in the Long Run”, *Journal of Monetary Economics*, forthcoming.

Mian, Atif and Amir Sufi, 2009, “The Consequences of Mortgage Credit Expansion: Evidence from the U.S. Mortgage Default Crisis”, *Quarterly Journal of Economics*, 124(4), 1449–1496.

Mian, Atif and Amir Sufi, 2011, “House Prices, Home Equity—Based Borrowing, and the US Household Leverage Crisis”, *American Economic Review*, 101(5), 2132–2156.

Pietrunti, Mario and Federico M. Signoretti, 2018, “Monetary Policy and the Mortgage Market: a New-Keynesian Perspective”, mimeo.
Richter, Björn, Moritz Schularick and Ilhyock Shim, 2019, “The Costs of Macroprudential Policy”, *Journal of International Economics*, 118(C), 263–282.

Rubio, Margarita and Yao Fang, 2018 “Macropuendental Policies in a Low Interest-Rate Environment”, *Journal of Money, Credit and Banking*, forthcoming.

Sajedi, Rana and Gregory Thwaites, 2016, “Why Are Real Interest Rates So Low? The Role of the Relative Price of Investment Goods”, *IMF Economic Review* 64(4), 635–659

Schmitt-Grohe, Stefanie and Martin Uribe, 2007, “Optimal Simple and Implementable Monetary and Fiscal Rules,” *Journal of Monetary Economics*, 54, 1702–1725.

Smets, Frank and Rafael Wouters, 2007, “Shocks and Frictions in US Business Cycles: A Bayesian DSGE Approach,” *American Economic Review*, 97(3), 586–606.

Walentin, Karl, 2014, “Housing Collateral and the Monetary Transmission Mechanism”, *Scandinavian Journal of Economics*, 116(3), 635–668.
## 8 Tables and Figures

Table 1: Calibrated structural parameters.

| Description                                                                 | Symbol | Value   |
|----------------------------------------------------------------------------|--------|---------|
| Capital share in the goods production function                            | $\mu_c$ | 0.35    |
| Capital share in the housing production function                          | $\mu_h$ | 0.10    |
| Land share in the housing production function                             | $\mu_l$ | 0.10    |
| Intermediate share in the housing production function                     | $\mu_o$ | 0.10    |
| Gross markup in prices and wages                                          | $X, X_{wh}, X_{wc}$ | 1.15    |
| Housing depreciation                                                       | $\delta_h$ | 0.005   |
| Capital depreciation, non-housing sector                                  | $\delta_{kc}$ | 0.025   |
| Capital depreciation, housing sector                                       | $\delta_{kh}$ | 0.03    |
| Taylor rule coefficient on inflation                                      | $r_\pi$ | 2.00    |
| Taylor rule coefficient on output gap                                     | $r_{\Delta \nu}$ | 0.50    |
| Taylor rule coefficient on output                                          | $r_y$ | 0.25    |
| Taylor rule smoothing                                                     | $\rho_R$ | 0.75    |
| Calvo price rigidity                                                      | $\theta_\pi$ | 0.9200  |
| Calvo wage rigidity, non-housing sector                                  | $\theta_{wc}$ | 0.7920  |
| Calvo wage rigidity, housing sector                                       | $\theta_{wh}$ | 0.9118  |
| Price indexation                                                          | $\nu_\pi$ | 0.6911  |
| Wage indexation, non-housing sector                                       | $\nu_{wc}$, | 0.08301 |
| Wage indexation, housing sector                                            | $\nu_{wh}$ | 0.41186 |
| Share of patient hhs                                                      | $\alpha$ | 0.67    |
| Consumption habit                                                         | $\varepsilon_P, \varepsilon_I$ | 0.70    |
| Capital utilization cost                                                  | $\zeta$ | 0.70    |
| Investment adjustment costs non-housing sector                            | $S_c'$ | 5.316   |
| Investment adjustment costs, housing sector                              | $S_h'$ | 7.485   |
| Inverse Frisch elasticity, patient hhs                                    | $\eta_P$ | 0.5238  |
| Inverse Frisch elasticity, impatient hhs                                  | $\eta_I$ | 0.5060  |
| Sectorial labor mobility, patient hhs                                     | $\xi_P$ | 0.6833  |
| Sectorial labor mobility, impatient hhs                                   | $\xi_I$ | 0.9654  |
| Amortization rate on hhs loans                                            | $\kappa$ | 0.0075  |
| Share refinancing                                                         | $\Phi$ | 0.3     |
| Housing preference weight, patient hhs                                    | $j_P$ | 0.1235  |
| Housing preference weight, impatient hhs                                   | $j_I$ | 0.2316  |
| Housing transaction costs                                                 | $\phi_h$ | 10      |
### Table 2: Parameters that drive the change in indebtedness.

| Parameter | Moment | 1990's | 2010's |
|-----------|--------|--------|--------|
|           | Value  | Target | Value  | Target  |
| $\beta_P$ | Real rate | 0.9925 | 3% | 0.99875 | 0.5% |
| $\theta_{LTV}$ | LTV | 0.75 | 75% | 0.85 | 85% |
| $\bar{\pi}$ | Inflation rate | 0.005 | 2% | 0.00375 | 1.5% |
| $\gamma$ | HEW fraction | 0.015 | - | 0.02095 | LTI=433% |

### Table 3: Long-run equilibrium in low debt vs. high debt economy (in percent)

|                  | 1990’s | 2010’s |
|------------------|--------|--------|
|                  | LTV | LTI | DSTI | LTV | LTI | DSTI |
| LTI borrowers     | 217 | 217 | 217 | 433 | 433 | 433 |
| DSTI (after tax) borrowers | 14.2 | 14.2 | 14.2 | 19.1 | 19.1 | 19.1 |
| Interest (after tax)/income of borrowers | 7.67 | 7.67 | 7.66 | 6.08 | 6.08 | 6.07 |
| Non-residential investment /GDP | 17.1 | 17.2 | 17.2 | 20.9 | 20.9 | 21.0 |
| Residential investment /GDP | 3.0 | 2.6 | 2.6 | 5.2 | 4.4 | 4.4 |
| House prices (%Δ from 1990’s to 2010’) | 36.5 | 34.4 | 34.4 |

### Table 4: Steady state effects of MPP in the two indebtedness regimes (percent change).

|                  | 1990’s | 2010’s |
|------------------|--------|--------|
|                  | LTV | LTI | MID | LTI | DSTI | LTV | LTI | MID | LTI | DSTI |
| Aggregate LTI    | -10.2 | -10.2 | -10.2 | -10.2 | -10.2 | -10.2 | -10.2 | -10.2 | -10.2 |
| Output           | -0.23 | -0.27 | -0.34 | -0.34 | -0.32 | -0.37 | -0.56 | -0.56 |
| Consumption      | -0.11 | -0.12 | -0.33 | -0.34 | -0.07 | -0.08 | -0.56 | -0.56 |
| Non-residential investment | -0.17 | -0.20 | -0.34 | -0.34 | -0.20 | -0.24 | -0.56 | -0.56 |
| Residential investment | -2.72 | -3.48 | -0.37 | -0.37 | -3.05 | -3.56 | -0.58 | -0.58 |
| House prices     | -1.06 | -1.37 | -0.04 | -0.05 | -1.21 | -1.41 | -0.07 | -0.07 |
| DSTI (after tax) of borrowers | -10.2 | 6.24 | -10.2 | -10.2 | 2.09 | 10.2 | -10.2 | -10.2 |
| Interest/income (after tax) of borrowers | -10.2 | 20.2 | -10.2 | -10.2 | 28.32 | -10.2 | -10.2 | -10.2 |
| Consumption of borrowers | 1.07 | 1.20 | 0.37 | 0.38 | 0.98 | 1.11 | -0.33 | -0.33 |
| Housing of borrowers | -7.03 | -9.21 | 0.42 | 0.42 | -7.94 | -9.30 | -0.27 | -0.27 |
| Hours worked of borrowers | -1.05 | -1.23 | -1.31 | -1.31 | -1.11 | -1.27 | -1.79 | -1.79 |
| Income of borrowers | -0.26 | -0.32 | -0.34 | -0.34 | -0.38 | -0.44 | -0.56 | -0.56 |

Note: In the high debt environment, mortgage interest deductibility (MID) is completely removed. In the low debt environment, MID is reduced to 6.35% to obtain the same reduction in LTI.
Figure 1: International Evidence

Households debt-to-income ratio

Real mortgage rate

House prices

Residential investment

Annual rate of inflation

- Sweden
- Norway
- Denmark
- UK
- USA
- Australia
- Canada
Figure 2: LTV in Sweden
Figure 3: LTI and LTV in Sweden. Source: Swedish FSA Mortgage Survey, average 2015-2017
Figure 4: Aggregate effects of LTV tightening under alternative household leverage assumptions.

Panel A: Monetary policy constrained by ZLB for 8 quarters

Panel B: Monetary policy unconstrained
Figure 5: Disaggregate effects of permanent LTV tightening in a liquidity trap under alternative household leverage assumptions

1. Aggregate Consumption
2. Savers Consumption
3. Borrowers Consumption
4. Non-residential Investment
5. Residential Investment
6. Housing Price
7. Savers Housing
8. Borrowers Housing
9. DTI of borrowers (Perc. Points)
Figure 6: Effects of a contractionary monetary policy shock under alternative household leverage assumptions in the LTV model
Figure 7: Aggregate Effects of LTV, LTI, DSTI and MID tightening in a liquidity trap under high indebtedness
Figure 8: Disaggregate effects of LTV, LTI, DSTI and MID tightening in a liquidity trap under high indebtedness
The upper panels show unconditional welfare. Welfare is computed using second order approximation of the model and using the method of Hebden, Lindé and Svensson (2010) to handle the ZLB. Total welfare refers to the following expression, \( W_S = (1 - \beta_P) W_p + (1 - \beta_I) W_I \), following Lambertini et al. (2013). The lower panels show the distribution of the outcome of variables for an LTV of 0.7.
A Technical appendix

In this appendix, we first provide a complete description of the simple two-period model and macroprudential tightening in that model. We then present additional details of the quantitative model, including all the first order conditions.

A.1 Two-period economy

Agents solve the following problem

\[
\text{Max } \sum_{t=1}^{2} \beta^{t-1} \left( c_t + \log (h_t) - \frac{n_t^2}{2} \right) \quad \text{s.t.:}
\]

\[
st : c_t + q_t (h_t - h_{t-1}) = n_t - Rb_{t-1} + b_t
\]

\[
b_1 \leq \bar{B}_1
\]

\[
b_0 = 0, h_0 = \bar{H}, b_2 = \theta
\]

where we consider two set-ups with different specifications for the borrowing limit, \( \bar{B}_t \):

\[
LTV : \bar{B}_1 = \theta q_1 h_1
\]

or

\[
LTI : \bar{B}_1 = \theta n_1
\]

The problem can be rewritten in a more compact form:

\[
\text{Max } \sum_{t=1}^{2} \beta^{t-1} \left( n_t - Rb_{t-1} + b_t - q_t (h_t - h_{t-1}) + \log (h_t) - \frac{n_t^2}{2} \right) \quad \text{s.t.:}
\]

\[
LTV : b_1 \leq \theta q_1 h_1 \text{ or } LTI : b_1 \leq \theta n_1
\]

with the following first order conditions:

\[
b_t : 1 = \beta R + \mu
\]

LTV constraint

\[
h_1 : q_1 = \frac{1}{h_1} + \beta q_2 + \mu \theta q_1
\]

\[
h_2 : q_2 = \frac{1}{h_2}
\]

\[
n_t : n_t = 1, \ t = 1, 2
\]
LTI constraint

\[ h_1 : q_1 = \frac{1}{h_1} + \beta q_2 \]
\[ h_2 : q_2 = \frac{1}{h_2} \]
\[ n_t : n_1 = 1 + \mu \theta \]

A.1.1 Equilibrium

Market clearing condition

\[ h_t = \bar{H}, \ t = 1, 2 \]

LTV limit. The behavior of house prices and debt in equilibrium is described by:

\[ q_1 = \frac{1}{\bar{H}} + \beta q_2 + (1 - \beta R) \theta q_1 \]
\[ = \frac{(1 + \beta) \frac{1}{1 - (1 - \beta R) \theta \bar{H}}}{1 - (1 - \beta R) \theta \bar{H}} \]
\[ q_2 = \frac{1}{\bar{H}} \]
\[ q_1 - q_2 = \frac{1}{\bar{H}} \left( \frac{\beta + (1 - \beta R) \theta}{1 - (1 - \beta R) \theta} \right) > 0 \]

\[ b_2 = \theta \]
\[ b_1 = \theta q_1 \bar{H} = \frac{\theta (1 + \beta)}{1 - (1 - \beta R) \theta} \]
\[ b_1 - b_2 = \theta \left( \frac{\beta + (1 - \beta R) \theta}{1 - (1 - \beta R) \theta} \right) > 0 \]

Consumption is determined by the following budget constraints:

\[ c_2 = 1 - Rb_1 + b_2 = 1 - R \frac{\theta (1 + \beta)}{1 - (1 - \beta R) \theta} + \theta \]
\[ = 1 - R \frac{\theta (1 + \beta)}{1 - \mu \theta} + \theta \]
\[ = \frac{(1 - \mu \theta) (1 + \theta) - R \theta (1 + \beta)}{1 - \mu \theta} \]
\[ = \frac{1 - \theta^2 \mu - R \theta}{1 - \mu \theta}, \]
\[ c_1 = 1 + b_1 = 1 + \frac{\theta (1 + \beta)}{1 - (1 - \beta R) \theta}. \]
To make sure that consumption in the second period is non-negative, the following condition needs to be satisfied:

\[
\frac{1 - \theta^2 \mu - R\theta}{1 - \mu \theta} > 0.
\]

Given our assumption \( \theta < \beta < 1 \), the denominator of the expression above is always positive and so is the numerator:

\[
1 - \theta^2 \mu - R\theta = \\
1 - \theta^2 (1 - R\beta) - R\theta > \\
1 - \theta^2 (1 - R\beta) - R\beta > \\
(1 - R\beta) (1 - \theta^2) > 0
\]

**LTI limit.** In the LTI economy, the behavior of house prices in equilibrium is simply described by:

\[
q_1 = \frac{1}{H} + \beta q_2 \\
q_2 = \frac{1}{H}
\]

Differently from before, labor supply is higher in the first period:

\[
n_1 = 1 + \theta (1 - \beta R) \\
n_2 = 1
\]

Debt

\[
b_1 = \theta (1 + \theta (1 - \beta R)) \\
b_2 = \theta
\]

As before, consumption can determined by the intraperiod budget constraints:

\[
c_2 = n_2 - Rb_1 + b_2 = 1 - R\theta (1 + \theta (1 - \beta R)) + \theta \\
= 1 + \theta - R\theta (1 + \theta \mu) \\
= \theta (1 - R\theta \mu) + (1 - R\theta) \\
c_1 = n_1 + b_1 = (1 + \theta) (1 + \theta (1 - \beta R))
\]

Note that also in this case our assumption \( \theta < \beta < 1 \) is sufficient to ensure a positive consumption level in the second period since \( \theta < \beta < \frac{1}{R} < \frac{1}{R\mu} \).
A.1.2 Macroprudential experiment

We start from a situation where debt in period 2 in the two economies is equal:

\[ b_{2, LTI} = b_{2, LTV} = \theta \]

and we want to cut \( b_2 \) in both economies by a fraction \( x < 1 \)

\[ \theta' = x\theta \]

The cut is permanent and it affects also \( b_1 \).

**LTV contraction** The behavior of house prices in equilibrium is described by:

\[ q'_1 = \frac{(1 + \beta)}{1 - (1 - \beta R) \theta x H} \]
\[ q'_2 = \frac{1}{H} \]

\[ b'_1 = \theta x q'_1 H = \frac{x\theta (1 + \beta)}{(1 - (1 - \beta R) \theta x)} \]

Consumption

\[ c'_2 = 1 - Rb'_1 + b'_2 = 1 - R \frac{x\theta (1 + \beta)}{(1 - (1 - \beta R) \theta x)} + x\theta \]
\[ c'_1 = 1 + b'_1 = 1 + \frac{x\theta (1 + \beta)}{(1 - (1 - \beta R) x\theta)} \]

**LTI Contraction**

\[ q'_1 = \frac{1}{H} + \beta q'_2 \]
\[ q'_2 = \frac{1}{H} \]

Labor supply

\[ n'_1 = 1 + x\theta (1 - \beta R) \]
\[ n'_2 = 1 \]
Debt

\[ b'_1 = \theta x (1 + x\theta (1 - \beta R)) \]
\[ b'_2 = x\theta \]

Consumption

\[ c'_2 = 1 - Rb'_1 + b_2 = 1 - R\theta x (1 + x\theta (1 - \beta R)) + x\theta \]
\[ c'_1 = (1 + \theta x (1 - \beta R)) (1 + x\theta) \]

A.1.3 Comparison between the two economies

In the LTV economy

\[ \Delta b_{1,LT} = b'_{1,LT} - b_{1,LT} \]
\[ = \theta (1 + \beta) \left[ \frac{x}{(1 - (1 - \beta R) x\theta)} - \frac{1}{(1 - (1 - \beta R) \theta)} \right] \]
\[ = \theta (1 + \beta) \left[ \frac{x}{(1 - (1 - \beta R) x\theta)} - \frac{1}{(1 - (1 - \beta R) \theta)} \right] \]
\[ = \theta (1 + \beta) \left[ \frac{(1 - (1 - \beta R) \theta) x - (1 - (1 - \beta R) x\theta)}{(1 - (1 - \beta R) x\theta)(1 - (1 - \beta R) \theta)} \right] \]
\[ = -\frac{\theta (1 + \beta) (1 - x)}{(1 - \mu x\theta) (1 - \mu \theta)} \]

where we used \( \mu = (1 - \beta R) \). In the LTI economy:

\[ \Delta b_{1,LTI} = b'_{1,LTI} - b_{1,LTI} = \theta x (1 + \theta x (1 - \beta R)) - \theta (1 + \theta (1 - \beta R)) \]
\[ = \theta x (1 + \theta x\mu) - \theta (1 + \theta \mu) \]
\[ = -\theta (1 - x) (\theta \mu (1 + x) + 1) \]
\[ \Delta n_{1,LTI} = -(1 - x) \theta \mu \]
\[ \Delta b_{1,LTI} + \Delta n_{1,LTI} = -\theta (1 - x) (\mu (\theta (1 + x) + 1) + 1) \]
Now, let us compare the reduction in debt in both economies

\[ \Delta b_{\text{LTV}} - \Delta b_{\text{LTI}} = \left[ \frac{\theta (1 + \beta)(1 - x)}{(1 - \mu x \theta) (1 - \mu \theta)} - \theta (1 - x) (\mu \theta (1 + x) + 1) \right] \]

\[ = - \left[ \frac{\theta (1 + \beta)(1 - x) - \theta (1 - x) (\mu \theta (1 + x) + 1) (1 - \mu x \theta) (1 - \mu \theta)}{(1 - \mu x \theta) (1 - \mu \theta)} \right] \]

\[ = -(1 - x) \theta \left[ \frac{\beta + 1 - (\mu \theta (1 + x) + 1) (1 - \mu x \theta) (1 - \mu \theta)}{(1 - \mu x \theta) (1 - \mu \theta)} \right] \]

\[ < 0 \]

This implies that debt decreases more in the LTV economy, i.e. that house prices respond more than labor supply to the policy change. In terms of consumption, in period 2:

\[ \Delta c_{2,\text{LTV}} = -R \Delta b_{1,\text{LTV}} + \Delta b_{2,\text{LTV}} \]

\[ \Delta c_{2,\text{LTI}} = -R \Delta b_{1,\text{LTI}} + \Delta b_{2,\text{LTI}} \]

\[ \Delta c_{2,\text{LTV}} - \Delta c_{2,\text{LTI}} = R (\Delta b_{1,\text{LTI}} - \Delta b_{1,\text{LTV}}) > 0 \]

In period 1

\[ \Delta c_{1,\text{LTV}} = \Delta b_{1,\text{LTV}} \]

\[ \Delta c_{1,\text{LTI}} = \Delta n_{1,\text{LTI}} + \Delta b_{1,\text{LTI}} = -\theta (1 - x) (\theta \mu (1 + x) + 1 + \mu) \]

We can now compute the difference between the drop in consumption in the two models:

\[ \Delta c_{1,\text{LTV}} - \Delta c_{1,\text{LTI}} = \Delta b_{\text{LTV}} - \Delta b_{\text{LTI}} - \Delta n_{1,\text{LTI}} \]

\[ = - \left[ \frac{\theta (1 + \beta)(1 - x)}{(1 - \mu x \theta) (1 - \mu \theta)} - \theta (1 - x) (\mu (\theta (1 + x) + 1) + 1) \right] \]

\[ = - \left[ \frac{\theta (1 + \beta)(1 - x) - \theta (1 - x) (\mu (\theta (1 + x) + 1) + 1) (1 - \mu x \theta) (1 - \mu \theta)}{(1 - \mu x \theta) (1 - \mu \theta)} \right] \]

\[ = -(1 - x) \theta \left[ \frac{\beta + 1 - (\mu (\theta (1 + x) + 1) + 1) (1 - \mu x \theta) (1 - \mu \theta)}{(1 - \mu x \theta) (1 - \mu \theta)} \right] \]

\[ = -(1 - x) \theta \left[ \frac{\theta^2 \mu^2 ((1 - \theta \mu) x (1 + x + \frac{1}{\theta}) + 1) + \beta - \mu (1 - \theta \mu)}{(1 - \mu x \theta) (1 - \mu \theta)} \right] < 0 \]
where the last inequality follows from

\[
\beta > \frac{1}{1+R} \\
\beta + R\beta > 1 \\
\beta > \mu
\]

That is, the drop in consumption is higher in the LTV case, Q.E.D.

### A.2 DSGE model

This section reports all first-order conditions of the quantitative DSGE model. The superscript prime refers to impatient households.

\[
\begin{align*}
ct + \frac{i_{ct}}{a_{k,t}} + n_{ht} + q_t h_t + l_t &= w_{ct} n_{ct} + w_{ht} n_{ht} + R_{ct} z_{ct} k_{ct-1} + R_{ht} z_{ht} k_{ht-1} \\
&+ (r_{t-1}^M + \kappa) \frac{d_{t-1}}{\pi_t} + \mu_t q_t IH_t \\
&+ q_t (1 - \delta_h) h_{t-1} - \frac{\phi_h}{2} \left( \frac{h_t}{h_{t-1}} - 1 \right)^2 q_t h_{t-1} \\
&+ \left( 1 - \frac{1}{X_t} \right) Y_t - a\left(z_{ct}\right) p_t^{kc} k_{ct-1}^{kc} a_{k,t} - a\left(z_{ht}\right) p_t^{kh} k_{ht-1} - T_t 
\end{align*}
\]

\[Div_t = \left( 1 - \frac{1}{X_t} \right) Y_t\]

\[
u_{c,t} q_t \left[ 1 + \Phi_h \left( \frac{h_t}{h_{t-1}} - 1 \right) \right] = a_{z,t} a_{j,t} \frac{j_t}{h_t}
\]

\[+ u_{c,t+1} \beta q_{t+1} (1 - \delta) \]

\[+ \left[ \Phi_h^2 u_{c,t+1} \beta q_{t+1} \left( \frac{h_{t+1}}{h_t} - 1 \right)^2 \right]\]

\[u_{h,t} = \frac{j_t}{h_t}\]
\[ u_{ct} = \beta u_{ct+1} R_t / \pi_{t+1} \]  

(19)

\[ \lambda_t^{dp} + r_t^M \lambda_t^{rP} = \frac{\beta}{\pi_{t+1}} \left[ u_{ct+1} (r_t^M + \kappa) + (1 - \kappa) \left( \lambda_t^{dp} + \lambda_t^{rP} [(1 - \Phi) r_t^M + \Phi r_t^F] \right) \right] \]  

(20)

\[ p_{t}^{kc} u_{ct} = \beta u_{ct+1} \left[ R_{ct+1} z_{ct+1} + p_{t+1}^{kc} \left( 1 - \delta_{kc} - a \frac{(z_{ct+1})}{a_{kt+1}} \right) \right] \]  

(21)

\[ u_{c,t} = \frac{\beta u_{c,t+1} \left( R_{ht+1} z_{ht+1} + p_{t+1}^{kh} \left( 1 - \delta_{kh} - a (z_{ht+1}) \right) \right)}{p_{t}^{kh}} \]  

(22)

\[ 1 = p_{t}^{kc} a_{kt} \left[ 1 - S_{c,t} - S_{c,t}' \frac{\Gamma_{AK} i_{ct}}{i_{ct-1}} \right] + \beta \frac{u_{c,t+1}^{kc}}{u_{c,t}} p_{t+1}^{kc} a_{kt} S_{c,t}' \left( \frac{i_{t+1}}{i_{ct}} \right)^2 \]  

(23)

\[ 1 = p_{t}^{kh} \left[ 1 - S_{h,t} - S_{h,t}' \frac{i_{ht}}{i_{ht-1}} \right] + \beta \frac{u_{c,t+1}^{kh}}{u_{c,t}} p_{t+1}^{kh} S_{h,t}' \left( \frac{i_{ht+1}}{i_{ht}} \right)^2 \]  

(24)

\[ a_t a_{z,t} (n_{c,t}^{1+\xi} + n_{h,t}^{1+\xi} \frac{n_{c,t}^{1+\xi}}{n_{c,t}^{1+\xi} + n_{h,t}^{1+\xi}}) = u_{et} \frac{w_{et}}{X_{cet}} \]  

(25)

\[ a_t a_{z,t} (n_{c,t}^{1+\xi} + n_{h,t}^{1+\xi} \frac{n_{c,t}^{1+\xi}}{n_{c,t}^{1+\xi} + n_{h,t}^{1+\xi}}) = u_{et} \frac{w_{ht}}{X_{whet}} \]  

(26)

\[ d_t = \frac{(1 - \kappa)}{\pi_t} d_{t-1} + l_t \]  

(27)

\[ r_t^{M} d_t = (1 - \Phi) (1 - \kappa) d_{t-1} \frac{r_{t-1}^{M}}{\pi_t} + \left[ l_t + \Phi (1 - \kappa) \frac{d_{t-1}}{\pi_t} \right] r_t^{F} \]  

(28)

\[ u_{ct} = \lambda_t^{dp} + \lambda_t^{rP} r_t^{F} \]  

(29)
\[ \lambda^t_P = \frac{\beta}{\pi_{t+1}} \left( u_{ct+1} + (1 - \Phi)(1 - \kappa)\lambda^t_{pt} \right) \quad (30) \]

\[ \text{transfer}'_t = T'_t = \frac{r^M_{t-1} \tau_{t-1} d_{t-1}}{\pi_t} \quad (31) \]

\[ T_t = 0 \quad (32) \]

\[ S'_e = \frac{1}{2} \sqrt{S'_c} \left\{ \exp \left[ \frac{h'_t}{h'_{t-1}} \right] - 1 \right\} - \exp \left[ -\sqrt{S'_c} \left( \frac{i_{ct}}{i_{ct-1}} - 1 \right) \right] \]

\[ S'_h = \frac{1}{2} \sqrt{S'_h} \left\{ \exp \left[ \frac{h'_t}{h'_{t-1}} \right] - 1 \right\} - \exp \left[ -\sqrt{S'_h} \left( \frac{i_{ht}}{i_{ht-1}} - 1 \right) \right] \]

\[ c'_t + q h'_t + [r^M_{t-1} + \kappa] \frac{d_{t-1}}{1\pi_t} + \left[ \Phi^{\beta'} \left( \frac{h'_t}{h'_{t-1}} \right) - 1 \right]^2 q h'_t \]

\[ = w'_{ct} n'_{ct} + w'_{ht} n'_{ht} + q_t (1 - \delta_h) \frac{h'_{t-1}}{1} + l_t + \tau r^M_{t-1} \frac{d_{t-1}}{1\pi_t} - T'_t \quad (33) \]

\[ u'_{ct} q_t \left[ 1 + \Phi^{\beta'} \left( \frac{h'_t}{h'_{t-1}} \right) \right] = a_{zt} a'_{zt} \frac{j'_t}{h'_{t-1}} \]

\[ + \left[ \Phi^{\beta'} u'_{ct+1} \beta' q_{t+1} (1 - \delta) \right] + \left[ \frac{\Phi^{\beta'}}{2} u'_{ct+1} \beta' q_{t+1} \left( \left( \frac{h'_{t+1}}{h'_t} \right)^2 - 1 \right) \right] + \lambda^{lt} m^{lt} q_t - \lambda^{ltv} \lambda^{ltv} q_{t+1} (1 - \delta_h)(m^{ltv}_{t+1} - \gamma) \quad (34) \]

\[ l_t = m^{lt} q_t [h'_t - (1 - \delta_h) h'_{t-1}] + \gamma [q_t (1 - \delta_h) h'_{t-1} - (1 - \kappa) \frac{d_{t-1}}{\pi_t}] \quad (35) \]

\[ \lambda'_{dp} + \lambda'_{tP} r^M_t = \frac{\beta'}{\pi_{t+1}} \left\{ u'_{ct} (r^M_t (1 - \tau_{t+1}) + \kappa) + \lambda^{ltv}_{t+1} \gamma (1 - \kappa) + (1 - \kappa) \left[ \lambda^{dp}_{t+1} + \lambda^{tP}_{t+1} (1 - \Phi) r^M_t + \Phi^{tF}_{t+1} \right] \right\} \quad (36) \]
\begin{align*}
& a_t a_{z,t} (n_{c,t}^{\epsilon'_{t}^1} + n_{h,t}^{\epsilon'_{t}^1})^{\frac{\mu_c - \epsilon_{t}^1}{1 + \epsilon_{t}^1}} n_{c,t}^{\epsilon'_{t}^1} = u'_{c,t} \frac{w'_{ct}}{X_{wct}} \quad (37) \\
& a_t a_{z,t} (n_{c,t}^{\epsilon'_{t}^1} + n_{h,t}^{\epsilon'_{t}^1})^{\frac{\mu_c - \epsilon_{t}^1}{1 + \epsilon_{t}^1}} n_{h,t}^{\epsilon'_{t}^1} = u'_{c,t} \frac{w'_{ht}}{X_{wht}} \quad (38) \\
& u'_{c,t} - \lambda_t^{lt} - \lambda_t^{dP} - \lambda_t^{rP} r_t F = 0 \quad (39) \\
& \lambda_t^{rP} = \frac{\beta_t}{\pi_{t+1}} \left( u'_{c,t+1} (1 - \tau_{t+1}) + (1 - \Phi) (1 - \kappa) \lambda_{t+1}^{rP} \right) \quad (40) \\
& Y_t = a_{c,t}^{1 - \mu_c} \left( n_{c,t}^{\alpha_{t}^1} n_{c,t}^{\alpha_{t}^1 - \alpha} \right)^{1 - \mu_c} \left( z_{c,t} k_{c,t-1} \right)^{\mu_c} \quad (41) \\
& IH_t = \left( a_{h,t} (n_{h,t}^{\alpha_{t}^1} n_{h,t}^{\alpha_{t}^1 - \alpha}) \right)^{1 - \mu_h - \mu_b - \mu_l} (z_{h,t} k_{h,t-1})^{\mu_h} (\mu_b q_t IH_t)^{\mu_b} \quad (42) \\
& (1 - \mu_c) \alpha \frac{Y_t}{X_t n_{ct}} = w_{ct} \quad (43) \\
& (1 - \mu_c) (1 - \alpha) \frac{Y_t}{X_t n_{ct}'} = w'_{ct} \quad (44) \\
& (1 - \mu_h - \mu_b - \mu_l) \alpha \frac{q_t IH_t}{n_{h,t}} = w_{ht} \quad (45) \\
& (1 - \mu_h - \mu_b - \mu_l) (1 - \alpha) \frac{q_t IH_t}{n_{h,t}'} = w'_{ht} \quad (46) \\
& \frac{\mu_c}{X_t} \frac{Y_t}{k_{ct-1}} = R_{ct} z_{ct} \quad (47) \\
& \frac{\mu_h}{k_{ht-1}} = R_{ht} z_{ht} \quad (48)
\end{align*}
\[
\log(\pi_t - \pi^*) - \nu_\pi \log(\pi_{t-1} - \pi^*) = \beta \left( \log(\pi_{t+1} - \pi^*) - \nu_\pi \log(\pi_t - \pi^*) \right) \\
- \frac{(1 - \theta_\pi)(1 - \beta \theta_\pi)}{\theta_\pi} \log \left( \frac{X_t}{X_{ss}} \right) + \log(\epsilon_{p,t}) \quad (49)
\]

\[
R_t = \frac{R_{t-1}^r}{(1 - r_R)\bar{r}} \left( \frac{Y_t}{Y_{t-1}} \right)^{(1 - r_R)\gamma} \left( \log \left( \frac{1}{\beta} \right) + \pi_{ss} \right)^{1 - r_R} \frac{\epsilon_{mp,t}}{a_s} \quad (50)
\]

\[
h_t + h'_t = (1 - \delta)(h_{t-1}/1) + (1 - \delta)(h'_{t-1}/1) + IH_t \quad (51)
\]

\[
u_{ct} = a_{z,ct} \frac{1 - \epsilon}{1(1 - \beta \epsilon)} \left( \frac{1}{c_t - \epsilon c_{t-1}} - \frac{\beta \epsilon}{c_{t+1} - \epsilon c_t} \right) \quad (52)
\]

\[
u'_{ct} = a_{z,ct} \frac{1 - \epsilon'}{1(1 - \beta' \epsilon')} \left( \frac{1}{c'_{t} - \epsilon' c'_{t-1}} - \frac{\beta' \epsilon'}{c'_{t+1} - \epsilon' c'_t} \right) \quad (53)
\]

\[
\ln(w_{ct}) = \frac{1}{1 + \beta_1} \ln(w_{c,t-1}) + \left( \frac{1}{1 + \beta_1} \right)(\ln(w_{c,t-1}) + \ln\pi_{t+1} - \ln\pi_{ss}) \\
- \frac{(1 + \beta_1 t_{wc})}{1 + \beta_1} (\pi_t - \pi_{ss}) + \frac{t_{wc}}{1 + \beta_1} (\ln\pi_{t-1} - \ln\pi_{ss}) \\
- \frac{(1 - \theta_{wc})(1 - \beta_1 \theta_{wc})}{\theta_{wc}} (1 + \beta_1) \ln \left( \frac{X_{wc,t}}{X_{ss}} \right) \quad (54)
\]

\[
\ln(w'_{ct}) = \frac{1}{1 + \beta'} \ln(w'_{c,t-1}) + \left( \frac{1}{1 + \beta'} \right)(\ln(w'_{c,t-1}) + \ln\pi_{t+1} - \ln\pi_{ss}) \\
- \frac{(1 + \beta' t_{wc})}{1 + \beta'} (\pi_t - \pi_{ss}) + \frac{t_{wc}}{1 + \beta'} (\ln\pi_{t-1} - \ln\pi_{ss}) \\
- \frac{(1 - \theta_{wc})(1 - \beta' \theta_{wc})}{\theta_{wc}} (1 + \beta') \ln \left( \frac{X'_{wc,t}}{X_{ss}'} \right) \quad (55)
\]
\[ \ln(w_{ht}) = \frac{1}{1 + \beta} \ln(w_{h,t-1}) + \left(\frac{1}{1 + \beta}\right) \left(\ln(w_{h,t-1}) + \ln\pi_{t+1} - \ln\pi^{*}\right) \\
- \frac{(1 + \beta') \ln\pi_{t} - \pi^{*}}{1 + \beta'} + \frac{t_{wh}}{1 + \beta} \left(\ln\pi_{t-1} - \ln\pi^{*}\right) \\
- \frac{(1 - \theta_{wh})(1 - \beta') X_{wh,t}}{(1 + \beta')} \ln X_{w}^{*} \tag{56} \]

\[ \ln(w'_{ht}) = \frac{1}{1 + \beta'} \ln(w'_{h,t-1}) + \left(\frac{1}{1 + \beta'}\right) \left(\ln(w'_{h,t-1}) + \ln\pi_{t+1} - \ln\pi^{*}\right) \\
- \frac{(1 + \beta') \ln\pi_{t} - \pi^{*}}{1 + \beta'} + \frac{t_{wh}}{1 + \beta'} \left(\ln\pi_{t-1} - \ln\pi^{*}\right) \\
- \frac{(1 - \theta_{wh})(1 - \beta') X'_{wh,t}}{(1 + \beta')} \ln X_{w}^{*} \tag{57} \]

\[ \frac{R_{ct}}{p_{ct}^{\zeta_{kc}}} = \frac{\zeta_{kc}}{1 - \zeta_{kc}} z_{kc,t} + \left(1 - \frac{\zeta_{kc}}{1 - \zeta_{kc}}\right) \tag{58} \]

\[ \frac{R_{ht}}{p_{ht}^{\zeta_{kh}}} = \frac{\zeta_{kc}}{1 - \zeta_{kc}} z_{kh,t} + \left(1 - \frac{\zeta_{kc}}{1 - \zeta_{kc}}\right) \tag{59} \]

\[ k_{ct} = (1 - \delta_{kc}) k_{ct-1} + (1 - S_{ct}) i_{ct} \tag{60} \]

\[ k_{ht} = (1 - \delta_{kh}) k_{ht-1} + (1 - S_{ht}) i_{ht} \tag{61} \]

\[ data - Y_t = c_t + c_t' + k_{ct} - (1 - \delta_{kc}) k_{ct-1} + k_{ht} - (1 - \delta_{kh}) \left(k_{ht-1} + q_t IH_t\right) \]

\[ a(z_{ct+1}) = \left(\frac{1}{\beta} - (1 - \delta_{kc})\right) \]

\[ - \left(0.5 \frac{\zeta_{kc}}{1 - \zeta_{kc}} z_{ct}^2 + (1 - \frac{\zeta_{kc}}{1 - \zeta_{kc}}) z_{ct} + (0.5 \frac{\zeta_{kc}}{1 - \zeta_{kc}} - 1)\right) \tag{62} \]
\[ a(z_{ht+1}) = \left( \frac{1}{\beta} - (1 - \delta_{kh}) \right) \]
\[ - \left( 0.5 \frac{\zeta_{kc}}{1 - \zeta_{kc}} z_{kht}^2 + (1 - \frac{\zeta_{kc}}{1 - \zeta_{kc}}) z_{kht} + (0.5 \frac{\zeta_{kc}}{1 - \zeta_{kc}} - 1) \right) \] (63)

\[ S_{ct} = \frac{1}{2} \left\{ \exp \left[ \sqrt{S_{c}^{\eta}} \left( \frac{1}{\lambda_{ct-1}} - 1 \right) \right] \right. \]
\[ + \exp \left[ -\sqrt{S_{c}^{\eta}} \left( \frac{1}{\lambda_{ct-1}} - 1 \right) \right] - 2 \} \] (64)

\[ S_{ht} = \frac{1}{2} \left\{ \exp \left[ \sqrt{S_{h}^{\eta}} \left( \frac{1}{\lambda_{ht-1}} - 1 \right) \right] \right. \]
\[ + \exp \left[ -\sqrt{S_{h}^{\eta}} \left( \frac{1}{\lambda_{ht-1}} - 1 \right) \right] - 2 \} \] (65)

### A.2.1 Shocks

The stochastic process for the exogenous shocks in the model are described below. All innovations are denoted by the letter \( \varepsilon \), with a subscript specifying the type. The standard deviations of these innovations are denoted by \( \sigma \) with the corresponding subscript. The preference shocks are AR(1) processes:

\[ \log z_t = \rho_z \log z_{t-1} + \varepsilon_{z,t} \]
\[ \log j_{c,t} = \rho_j \log j_{c,t} + (1 - \rho_j) \log j_c + \varepsilon_{j,c,t} \quad c = \{ P, I \} \]
\[ \log v_t = \rho_v \log v_{t-1} + \varepsilon_{v,t} \]

Shocks to the LTV, LTI and DSTI requirements, \( \theta \), are very persistent AR(1) processes as described in the main text. Interest rate deductibility (\( \tau \)) follows an AR(2) process also described in the main text. As mentioned above, the cost-push shock, \( \varepsilon_{\pi,t} \), and the monetary policy shock, \( \varepsilon_{r,t} \), are i.i.d. Technology shocks are described in more detailed in the subsection below.
Technology shocks  We allow for three productivity processes: consumption goods, housing, and non-housing investment-specific productivity. The three processes are:

\[
\log a_{ct} = \rho_{AC} \log a_{ct-1} + \varepsilon_{ct}
\]

\[
\log a_{ht} = \rho_{AH} \log a_{ht-1} + \varepsilon_{ht}
\]

\[
\log a_{kt} = \rho_{AK} \log a_{kt-1} + \varepsilon_{kt}
\]

A.2.2 Definition of investment adjustment cost function and utilization cost

The investment adjustment costs can be expressed as:

\[
F(i_t, i_{t-1}) = \left(1 - \tilde{S}\left(\frac{i_t}{i_{t-1}}\right)\right) i_t \quad s = \{K, H\}
\]

where

\[
\tilde{S}(x) = \frac{1}{2} \left\{ \exp \left[ \sqrt{S''_s} (x - \Gamma_s) \right] + \exp \left[ -\sqrt{S''_s} (x - \Gamma_s) \right] - 2 \right\} \quad s = \{SK, SH\}
\]

The capital utilization cost function is (with the same parameter \(\zeta\) for both sectors)

\[
a(z_{st}) = R_s \left( \frac{\zeta}{2} \frac{1}{1 - \zeta} z_{st}^2 + \left(1 - \frac{\zeta}{1 - \zeta}\right) z_{st} + \left(\frac{1}{2} - \frac{\zeta}{1 - \zeta}\right) - 1 \right) \quad s = \{c, h\}
\]

A.2.3 Wage equations

The wage equations for each sector–household pair is:

\[
\omega_{c,P,t} - \tau_{wc} \log \pi_{t-1} = \beta_P (E_t \omega_{c,P,t+1} - \tau_{wc} \log \pi_t) - \varepsilon_{w,P,c} \log \left( \frac{X_{wct}}{X_{wc}} \right)
\]

\[
\omega_{c,I,t} - \tau_{wc} \log \pi_{t-1} = \beta_I (E_t \omega_{c,I,t+1} - \tau_{wc} \log \pi_t) - \varepsilon_{w,I,c} \log \left( \frac{X_{wct}}{X_{wc}} \right)
\]

\[
\omega_{h,P,t} - \tau_{wh} \log \pi_{t-1} = \beta_P (E_t \omega_{h,P,t+1} - \tau_{wh} \log \pi_t) - \varepsilon_{w,P,c} \log \left( \frac{X_{wht}}{X_{wh}} \right)
\]

\[
\omega_{h,I,t} - \tau_{wh} \log \pi_{t-1} = \beta_I (E_t \omega_{h,I,t+1} - \tau_{wh} \log \pi_t) - \varepsilon_{w,I,c} \log \left( \frac{X_{wht}}{X_{wh}} \right)
\]

where \(\omega_{it}\) denotes log nominal wage inflation, i.e. \(\omega_{it} = w_{it} - w_{it-1} + \pi_t\). \(\varepsilon_{wc}, \varepsilon_{wh}\) are defined below.
A.2.4 Definitions of various parameters

\[
\varepsilon_x = (1 - \theta_x) \left( 1 - \beta P \theta_x \right) / \theta_x
\]
\[
\varepsilon_{wPc} = (1 - \theta_{wc}) \left( 1 - \beta P \theta_{wc} \right) / \theta_{wc}
\]
\[
\varepsilon_{wIc} = (1 - \theta_{wc}) \left( 1 - \beta I \theta_{wc} \right) / \theta_{wc}
\]
\[
\varepsilon_{wPh} = (1 - \theta_{wh}) \left( 1 - \beta P \theta_{wh} \right) / \theta_{wh}
\]
\[
\varepsilon_{wIh} = (1 - \theta_{wh}) \left( 1 - \beta I \theta_{wh} \right) / \theta_{wh}
\]

A.3 Appendix Figures
Figure A.1: Effects of a contractionary monetary policy shock in models with different borrowing constraints
Figure A.2: Aggregate effects of MID removal across different borrowing constraints with unconstrained monetary policy and high indebtedness.