Reproductive profiles and risk of breast cancer subtypes: a multi-center case-only study

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Abstract

Background: Previous studies have shown that reproductive factors are differentially associated with breast cancer (BC) risk by subtypes. The aim of this study was to investigate associations between reproductive factors and BC subtypes, and whether these vary by age at diagnosis.

Methods: We used pooled data on tumor markers (estrogen and progesterone receptor, human epidermal growth factor receptor-2 (HER2)) and reproductive risk factors (parity, age at first-time pregnancy (FFTP) and age at menarche) from 28,095 patients with invasive BC from 34 studies participating in the Breast Cancer Association Consortium (BCAC). In a case-only analysis, we used logistic regression to assess associations between reproductive factors and BC subtype compared to luminal A tumors as a reference. The interaction between age and parity in BC subtype risk was also tested, across all ages and, because age was modeled non-linearly, specifically at ages 35, 55 and 75 years.

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**Results:** Parous women were more likely to be diagnosed with triple negative BC (TNBC) than with luminal A BC, irrespective of age (OR for parity = 1.38, 95% CI 1.16–1.65, \( p = 0.0004 \); \( p \) for interaction with age = 0.076). Parous women were also more likely to be diagnosed with luminal and non-luminal HER2-like BCs and this effect was slightly more pronounced at an early age (\( p \) for interaction with age = 0.037 and 0.030, respectively). For instance, women diagnosed at age 35 were 1.48 (CI 1.01–2.16) more likely to have luminal HER2-like BC than luminal A BC, while this association was not significant at age 75 (OR = 0.72, CI 0.45–1.14). While age at menarche was not significantly associated with BC subtype, increasing age at FFTP was non-linearly associated with TNBC relative to luminal A BC. An age at FFTP of 25 versus 20 years lowered the risk for TNBC (\( OR = 0.78, CI 0.70–0.88, p < 0.0001 \), but this effect was not apparent at a later FFTP.

**Conclusions:** Our main findings suggest that parity is associated with TNBC across all ages at BC diagnosis, whereas the association with luminal HER2-like BC was present only for early onset BC.

**Keywords:** Breast cancer subtype, Age at breast cancer diagnosis, Parity, Age at first full-time pregnancy, Age at menarche

**Background**

Worldwide, breast cancer (BC) is the most frequently diagnosed malignancy and leading cause of female cancer death [1]. Over the past decade, it has become evident that BC represents a heterogeneous disease, for which different subtypes can be distinguished based on the combination of tumor grade and the presence of hormone receptors, i.e., estrogen (ER), progesterone (PR) and human epidermal growth factor receptor-2 (HER2). Each BC subtype, including the luminal A-like, luminal B-like, luminal HER2-like, HER2-like and triple negative breast cancer (TNBC), presents with different age and risk factor distributions [2]. Analyses from the Breast Cancer Association Consortium (BCAC) showed, for instance, that nulliparity and a later age at first full-time pregnancy (FFTP) increase the risk of ER-positive BC, but not ER-negative BC [3].

Menarche and FFTP, and in particular their timing, may have diverse and complex effects on BC risk. It has been proposed that pregnancy induces the protective differentiation of mammary cells in the terminal duct lobular unit, which translates into long-term protection against BC [4–6]. Subsequent full-term pregnancies exert a similar but quantitatively much less important effect, which is a likely reflection of the protective differentiation of breast cells already induced by the FFTP [7]. The protective effect of an FFTP, however, is not apparent when the FFTP occurs after the age of 35 years [8, 9]. Although an FFTP offers long-term protection against BC, pregnancy is also associated with a transient increased BC risk postpartum, which could be due to pregnancy-related stimulation of pre-existent malignant clones [7, 10, 11]. In addition, the protective effect of the FFTP is found to be greater later in life [12], which according to a hypothesis by Russo et al. can be explained by the fact that several breast tumors are already initiated before the pregnancy, i.e., before the FFTP can induce its protective effect [4, 5].

Age at menarche has also been reported to influence BC risk differently depending on age. A late age at menarche is associated with a later onset of ovulatory cycles, and consequently with a decreased lifetime exposure to estrogen [13, 14]. For instance, per one year younger age at menarche, the associated BC risk increases by about 7% in women aged < 45 years, whereas the increase is about 4% in women aged 65 or older [15]. Also a short window of susceptibility, which is defined as the time between age at menarche and age at FFTP, lowers BC risk [16].

So far, few studies have examined how reproductive variables may be differentially associated with the risk of a specific BC subtype [2, 17]. The majority of these studies did not consider differential effects by age or were typically limited to cases developing BC at young age [18, 19], or to postmenopausal women only [20, 21]. Additional investigations involving women diagnosed with BC in all age categories, for which data on BC subtypes are available, are thus warranted. Therefore, we examined the association between parity and the risk of developing a specific BC subtype, and how this may differ according to age. Furthermore, we assessed the association of age at menarche (in nulliparous and parous women) and age at FFTP with the risk of being diagnosed with a specific BC subtype.

**Methods**

**BCAC cohorts, inclusion and exclusion criteria**

This analysis includes data from studies which participate in the BCAC and could provide information on BC risk factors, in particular parity (never versus ever), age at menarche and age at FFTP, and clinic-pathological information, in particular, grade, ER, PR and HER2 status of the tumor. Studies not providing any of these data were excluded. Three additional studies in the BCAC with information on BC subtypes were also not included: two studies included only ER-negative BC or patients
with TNBC (SKKDKFZS, NBCS); one study included only ER-positive/non-TNBC patients (PBCS). Overall, 34 of the 49 available studies in the BCAC provided data justifying inclusion in our study. These were composed of 10 population-based studies (8 case–control and 2 prospective cohort studies); 5 hospital-based case-control studies; and 19 studies of mixed design (all other studies). No information on previous in situ BC was available, though some studies excluded patients with a previous BC. Figure 1 displays the numbers of excluded and included patients. Patients with in situ breast cancer (N = 3932) and patients with bilateral BC (N = 2349) were excluded, since interpretation may be difficult in the case of coexistence of different BC phenotypes. Patients with BC diagnosed during pregnancy (N = 23) were excluded as well. Patients with insufficient information for assignment to BC subtype according to the 2011 St. Gallen criteria [22] (N = 18,612) were also excluded. Patients with BC before a first pregnancy were classified as nulliparous. Only BC patients with an unambiguously defined surrogate molecular BC status (N = 11,328), and with a known parity and age at BC diagnosis were included (Fig. 1).

**Tumor marker definitions and definition of breast cancer subtypes**

Definitions of ER, PR and HER2 status were not standardized across studies since most data were extracted from medical records (15 of 34 studies for ER and PR, and 6 of 34 studies for HER2). The 2011 St. Gallen criteria were used to define the five surrogate BC molecular phenotypes, using grade instead of Ki67 positivity [22]: (i) luminal A-like (ER-positive and/or PR-positive, HER2-negative, grade 1 or 2), (ii) luminal B-like (ER-positive and/or PR-positive, HER2-negative, grade 3), (iii) luminal HER2-like (ER-positive and/or PR-positive and HER2-positive), (iv) HER2-like (ER-negative and PR-negative, HER2-positive), and TN (ER-negative, PR-negative and HER2-negative) BC. How these internationally accepted definitions relate to other definitions of BC subtypes, as previously also applied in other BCAC publications, is depicted in Additional file 1: Table S1.

**Statistical methodology**

To evaluate the interaction between parity and age at BC diagnosis considering all BC subtypes simultaneously, a baseline-category logit model was applied to a five-category polytomous variable consisting of the five BC molecular subtypes, whereby luminal A-like breast cancer was considered the reference category. Interactions between parity and age at BC diagnosis in the probability of developing a specific BC subtype were tested by separate logistic regression models with binary outcome (1 for the BC subtype and 0 for luminal A-like BC as the reference subtype). For age at diagnosis, non-linear trends were explored by means of quadratic and cubic spline-based curves. Likelihood ratio testing was used for model selection of nested models and akaike information criterion (AIC) in the case of un-nested models. The best fit was obtained when modeling age non-linearly using cubic splines (five knots). Odds ratios
(OR) with 95% confidence intervals were estimated across all ages and, since age at diagnosis was modeled non-linearly, extrapolated from the non-linear fit at three different ages: 35 years (premenopausal age), 55 years (early postmenopausal age) and 75 years (late postmenopausal age). Hence, odds ratios for parity at selected ages are deduced from the curves generated for all patients at all ages, and are not based on subgroups of patients at selected ages.

To follow up on a significant interaction between age at diagnosis (continuous) and parity in luminal HER2-like and HER2-like BC, which only differ in their ER/PR status, we tested whether this interaction varied by combined ER/PR status (ER+/PR+, ER-/PR+, ER+/PR- = 1, ER-/PR- = 0) by means of a three-way interaction test (age at diagnosis × parity × ER/PR status). The test was conducted using a binary logistic regression model (HER2 status negative/positive) with HER2-negative as the reference group. Additionally, we assessed the association between parity and HER2-positive BC by age at diagnosis, using luminal A-like BC as the reference subtype.

Associations of age at menarche (in parous and nulliparous women) and FFTP (in parous women only) with BC subtypes were evaluated using logistic regression models. Again, luminal A-like BC was used as the reference subtype. Age at BC diagnosis was included in these models to correct for possible confounding. In parous women, age at FFTP was modelled according to the best fit, i.e., as a linear function for luminal B-like, luminal HER2-like and HER2-like BC, but as a quadratic function for TNBC.

To account for clustering of patients by study, a multi-level (or random-effects) model was used to provide unbiased standard errors and $p$ values [23]. All tests were two-sided and $p$ values smaller than 0.05 were considered significant. The analyses were performed using SAS software, version 9.2.

**Results**

Descriptive statistics for age at BC diagnosis, menopausal status, parity, age at menarche and age at FFTP stratified by molecular subtype are presented in Table 1. Additional information about these variables per individual study and for tumor size, nodal status, tumor grade and PR status can be found in Additional file 1: Tables S2-S4.

First, we assessed the association between parity and BC subtype compared to luminal A BC as a reference. This was done across all ages modeled non-linearly using cubic splines (see “Methods”), and because of the non-linear fit also extrapolated to specific ages. Table 2 includes specific estimates at 35, 55 and 75 years of age corresponding to premenopausal, early postmenopausal and late postmenopausal ages, whereas Additional file 1: Table S5 highlights estimates at 40, 50 and 60 years of age. A frequency table showing parity by BC subtype and specific age groups is provided as Additional file 1: Table S6. Parous women were more likely to develop TNBC compared to luminal A tumors ($OR = 1.38$, CI $1.16–1.65$, $p = 0.0004$, Table 2), but this association did not vary significantly by age ($p$ for interaction = 0.076). Graphical representation of these associations (Fig. 2a-d) nevertheless suggested that parous women were more likely to develop TNBC around age 55 years (Fig. 2d).

Compared to luminal A tumors, we did not observe a significant association between parity and luminal B-like BC across all ages ($OR = 0.90$, CI $0.77–1.05$, $p = 0.18$), nor at selected ages (Table 2 and Fig. 2a). For luminal HER2-like BC, there was also no significant association with parity across all ages ($OR = 1.04$, CI $0.88–1.24$, $p = 0.62$). We did detect, however, a weak but significant interaction between parity and age ($p$ for interaction = 0.037). Indeed, although confidence intervals were wide, parous women were slightly more likely to develop luminal HER2-like BC at age 35 ($OR = 1.48$, CI $1.01–2.16$, $p = 0.046$). This association was not significant in women aged 55 ($OR = 1.35$, CI $0.92–1.99$, $p = 0.13$) and was almost opposite in women aged 75 ($OR = 0.72$, CI $0.45–1.14$, $p = 0.16$, Table 2 and Fig. 2b). Associations between parity and HER2-like breast tumors were similar to those observed for luminal HER2-like BC (Table 2 and Fig. 2c). There was indeed a significant interaction between parity and age ($p$ for interaction = 0.030), but at specific ages ORs were not significant. Next, we combined luminal HER2-like and HER2-like BC and investigated whether parity may be associated with the likelihood of developing HER2+ BC (Table 3). A two-way interaction test between parity and age for HER2+ BC relative to luminal A BC was significant ($p$ for interaction = 0.003), but a three-way interaction test between parity, age and ER/PR status revealed that the interaction between parity and age does not differ by ER/PR status ($p$ for interaction = 0.49). Compared to luminal A BC, parous women were more likely to develop HER2+ BC at age 35 and 55 ($OR = 1.44$, CI $1.02–2.03$, $p = 0.037$ and $OR = 1.42$, CI $1.04–1.96$, $p = 0.029$, respectively), while an inverse association was observed at age 75 ($OR = 0.67$, CI $0.67–0.98$, $P = 0.041$, Table 3). Figure 3 visualizes how the association between parity and HER2+ BC differs by age.

Next, we assessed whether age at menarche affected the likelihood of being diagnosed with a specific BC subtype, while considering luminal A BC as the reference. The results when parity was considered a dichotomous variable are presented in Table 4, while the data
### Table 1: Distribution of reproductive risk factors according to the five surrogate BC subtypes

|                          | All BC subtypes | Luminal A-like N = 5914 (52.2%) | Luminal B-like N = 1553 (13.7%) | Luminal HER2-like N = 1509 (13.3%) | HER2-like N = 841 (74%) | TNBC N = 1511 (13.3%) |
|--------------------------|-----------------|---------------------------------|---------------------------------|-----------------------------------|-------------------------|------------------------|
| Age at breast cancer diagnosis, years | Mean/median 560/560 | 57.6/57.0 | 56.9/57.0 | 52.6/52.0 | 52.9/53.0 | 53.8/54.0 |
| <41                      | 1077 9.5%       | 328 5.5%   | 139 9.0%   | 247 16.4%  | 120 14.3%  | 243 16.1%   |
| 41–50                    | 2759 24.4%      | 1369 23.1% | 367 23.6%  | 455 30.2%  | 226 26.9%  | 342 22.6%   |
| 51–60                    | 51-60y 30.4%    | 1842 31.1% | 424 27.3%  | 406 26.9%  | 287 34.1%  | 481 31.8%   |
| >60                      | 4052 35.8%      | 2375 40.2% | 623 40.1%  | 401 26.6%  | 208 24.7%  | 445 29.5%   |
| Menopausal status        |                 |                   |                   |                   |                   |                       |
| Pre/peri                 | 3796 33.5%      | 1797 30.4%     | 505 32.5%     | 612 40.6%    | 316 37.6%   | 566 37.5%   |
| Post                     | 7075 62.5%      | 3886 65.7%     | 979 63.0%     | 835 55.3%    | 484 57.6%   | 891 59.0%   |
| Unknown                  | 457 4.0%        | 231 3.9%       | 69 4.4%       | 62 4.1%      | 41 4.9%     | 54 3.6%     |
| Parity (>/>=24 week-pregnancies, n) |               |                   |                   |                   |                   |                       |
| 0                        | 1746 15.4%      | 898 15.2%      | 266 17.1%     | 242 16.0%    | 135 16.1%   | 205 13.6%   |
| 1                        | 2062 18.2%      | 1052 17.8%     | 302 19.4%     | 285 18.9%    | 147 17.5%   | 276 18.3%   |
| 2                        | 4238 37.4%      | 2238 37.8%     | 547 35.2%     | 579 38.4%    | 302 35.9%   | 572 37.9%   |
| 3                        | 2102 18.6%      | 1117 18.9%     | 264 17.0%     | 270 17.9%    | 165 19.6%   | 286 18.9%   |
| 4 or more                | 1180 10.4%      | 609 10.3%      | 174 11.2%     | 133 8.8%     | 92 10.9%    | 172 11.4%   |
| Age at menarche, years   | Mean/median 132/130 | 13.2/13.0     | 13.2/13.0     | 13.3/13.0    | 13.4/13.0   | 13.2/13.0   |
| <12                      | 1369 13.3%      | 748 13.8%      | 186 13.5%     | 163 12.0%    | 82 11.1%    | 190 13.8%   |
| 12–13                    | 2104 20.5%      | 1066 19.7%     | 308 22.4%     | 289 21.3%    | 151 20.4%   | 290 21.0%   |
| 13–14                    | 2578 25.1%      | 1359 25.1%     | 318 23.1%     | 347 13.3%    | 188 25.3%   | 366 26.5%   |
| >14                      | 4227 41.1%      | 2249 41.5%     | 563 40.9%     | 561 33.3%    | 321 43.3%   | 533 38.7%   |
| Missing                  | 1050 492        | 178 149        | 99 132        |               |               |               |
| Age at FFTP, years       | Mean/median 25.1/240 | 25.0/240     | 25.3/250      | 25.6/250     | 25.6/250    | 24.7/240    |
| <20                      | 822 10.9%       | 453 11.2%      | 98 10.5%      | 89 9.2%      | 43 8.1%     | 139 13.2%   |
| 20–25                    | 2948 39.1%      | 1613 39.8%     | 353 38.0%     | 340 35.0%    | 200 37.8%   | 442 42.1%   |
| 25–30                    | 2417 32.1%      | 1292 31.8%     | 300 32.3%     | 344 35.4%    | 184 34.8%   | 297 28.3%   |
| >30                      | 1351 17.9%      | 699 17.2%      | 179 19.2%     | 198 20.4%    | 102 19.3%   | 173 16.5%   |
| Missing                  | 2044 1857       | 623 538        | 312 460       |               |               |               |

*BC* breast cancer, *HER2* human epidermal growth factor receptor-2, *TNBC* triple negative breast cancer, *Pre/peri* premenopausal/perimenopausal, *FFTP* first full-term pregnancy
Table 2: Association between parity (ever versus never) and BC subtypes for age overall and for specific ages (35, 55 and 75 years)

| Age at BC diagnosis | Odds ratio (95% CI) | P value | P value interaction parity × age |
|---------------------|---------------------|---------|---------------------------------|
| Luminal A-like      | All ages            | 1.00 (Ref) |        |                                  |
| Luminal B-like      | All ages            | 0.90 (0.77–1.05) | 0.18 |                                  |
| Luminal HER2-like   | All ages            | 1.04 (0.88–1.24) | 0.62 |                                  |
| HER2-like           | All ages            | 1.04 (0.83–1.29) | 0.73 |                                  |
| TNBC                | All ages            | 1.38 (1.16–1.65) | 0.0004 |                                |
| Luminal A-like      | At 35 years         | 1.00 (Ref) |        |                                  |
| Luminal B-like      | At 35 years         | 0.95 (0.64–1.43) | 0.82 |                                  |
| Luminal HER2-like   | At 35 years         | 1.48 (1.01–2.16) | 0.046 |                                |
| HER2-like           | At 35 years         | 1.35 (0.92–1.99) | 0.13 |                                  |
| TNBC                | At 35 years         | 1.38 (0.86–2.20) | 0.18 |                                  |
| Luminal A-like      | At 55 years         | 0.91 (0.65–1.26) | 0.55 |                                  |
| Luminal B-like      | At 55 years         | 0.89 (0.64–1.24) | 0.48 | 0.99                              |
| Luminal HER2-like   | At 55 years         | 1.35 (0.92–1.99) | 0.13 |                                  |
| HER2-like           | At 55 years         | 1.72 (0.45–1.14) | 0.16 | 0.037                             |
| TNBC                | At 55 years         | 1.38 (0.86–2.20) | 0.18 |                                  |
| Luminal A-like      | At 75 years         | 0.95 (0.64–1.43) | 0.82 |                                  |
| Luminal B-like      | At 75 years         | 0.89 (0.64–1.24) | 0.48 | 0.99                              |
| Luminal HER2-like   | At 75 years         | 1.14 (0.72–1.76) | 0.046 |                                |
| HER2-like           | At 75 years         | 1.24 (0.90–1.71) | 0.18 |                                  |
| TNBC                | At 75 years         | 1.24 (0.90–1.71) | 0.18 |                                  |

A baseline-category logits model was fitted with breast cancer (BC) subtype as a response variable taking luminal A BC as a reference category, and parity and age at diagnosis as a continuous variable as explanatory variables. Age was modeled non-linearly using cubic splines (five knots). The p value was 0.0149 for interaction effect between parity and age. A random intercept was introduced to account for clustering by study. Interactions between parity and age at BC diagnosis in the probability of developing a specific BC subtype were tested by logistic regression models with binary outcome (1 for the BC subtype and 0 for luminal A-like BC as the reference subtype). The interaction between age and parity in BC subtype risk was tested across all ages and, because age was modeled non-linearly, also specifically at age 35, 55 and 75 years. BC breast cancer, HER2 human epidermal growth factor receptor-2, TNBC triple negative breast cancer.

Discussion

Our analyses in pooled data on 11,328 patients with invasive BC showed that parity is associated with TNBC relative to luminal A disease, irrespective of age at diagnosis. A weak association between parity and luminal HER2-like BC on the other hand could only be observed when assessed at different ages. Furthermore, age at FFTP was non-linearly associated with TNBC.

In an earlier case-only study using BCAC data [3], we reported that parity is associated with a greater probability of being diagnosed with TNBC compared to ER+/HER- or PR+/HER- tumors. In the current analysis, this was confirmed, but relative to luminal A tumors. We did not observe that this association differed significantly according to age (p = 0.076), although the effects appeared slightly stronger with older age.

Also in line with previous results reported by the BCAC [3], parity was not associated with HER2-positive BC (both luminal HER2-like and HER2-like BC), while using luminal A-like (HER2-negative) BC as the reference. However, we did observe for the first time that this association differed significantly by age. Parous women diagnosed at 35 years of age were more likely to present with luminal HER2-like or HER2-like rather than luminal A BC (OR = 1.44), whereas parous women diagnosed after the menopause, at 75 years of age, were not (OR = 0.67). We hypothesize that pregnancy may promote HER2 positivity in BCs that are already subclinically present during pregnancy, but only become clinically apparent in the years following pregnancy. With older age, we found that parous women are less likely to have HER2-positive BC. Here, pregnancy could induce a protective effect against HER2-positive BCs. Phipps et al. also reported an association between late age at FFTP and increased risk of HER2-like BC, suggesting that there may only be a protective effect of pregnancy against HER2-positive BC when pregnancy precedes carcinogenesis [24].
Fig. 2 Association between parity and luminal B-like, luminal human epidermal growth factor receptor-2 (HER2)-like, HER2-like and triple-negative breast cancer by age at diagnosis. A binary logistic regression model was fitted considering every molecular subtype as the response variable, while considering luminal A-like breast cancer as a reference category, and parity and age as explanatory continuous variables. Age was modeled non-linearly using cubic splines (5 knots). Blue lines represent probabilities for parous women, green lines probabilities for nulliparous women.

- **a** Probability of luminal B-like subtype by parity.
- **b** Probability of Luminal HER2-like subtype by parity.
- **c** Probability of HER2-like subtype by parity.
- **d** Probability of triple negative breast cancer (TNBC) subtype by parity.

Table 3 Association between parity (ever versus never) and HER2+ BC at specific ages (age 35, 55 and 75 years)

| Age at BC diagnosis | Odds ratio (95% CI) | P value | P value interaction parity × age |
|---------------------|---------------------|---------|---------------------------------|
| Luminal A-like      |                     |         |                                 |
| At 35 years         | 1.00 (Ref.)         |         |                                 |
| At 55 years         | 1.00 (Ref.)         |         |                                 |
| At 75 years         | 1.00 (Ref.)         |         |                                 |
| HER2+ BC            |                     |         |                                 |
| At 35 years         | 1.44 (1.02–2.03)    | 0.037   |                                 |
| At 55 years         | 1.42 (1.04–1.96)    | 0.029   |                                 |
| At 75 years         | 0.67 (0.67–0.98)    | 0.041   | 0.003                           |

We combined luminal human epidermal growth factor receptor-2 (HER2)-like and HER2-like breast cancer (BC), and investigated whether parity may be associated with the risk of developing HER2+ BC. Interactions between parity and age at BC diagnosis on the probability to develop a specific BC subtype were tested by logistic regression models with binary outcome (1 for HER2+ BC and 0 for luminal A-like BC as the reference subtype) considering parity and age at diagnosis (as a continuous variable) as explanatory variables. A random intercept was introduced to account for clustering by study. The interaction between age and parity on HER2+ BC risk was tested, across all ages and, because age was modeled non-linearly, also specifically at age 35, 55 and 75 years.

Fig. 3 Association between parity and human epidermal growth factor receptor-2 (HER2) + breast cancer by age at diagnosis. A binary logistic regression model was fitted, considering HER2+ breast cancer as the response variable with luminal A-like breast cancer as the reference category. Age was modeled non-linearly using cubic splines (5 knots). Blue lines represent probabilities for parous women, green lines probabilities for nulliparous women.
Several studies already also reported that BC risk varies in the function of the time window between age at menarche and age at FFTP, with most studies suggesting that a short time interval is significantly and inversely associated with ER-positive BC [16, 18, 25, 26]. In the current study, we were able to provide a more detailed analysis of this effect. With respect to age at FFTP, we found that with an older age at FFTP, women were less likely to be diagnosed specifically with TNBC. Interestingly, this association seemed to be stronger for younger ages at FFTP (20–25 years) than for older ages at FFTP (30–35 years). On the other hand, age at menarche was not differentially associated with BC subtypes.

These findings should now be confirmed in large population-based studies. Indeed, the design of our case-only study, in which we calculated odds ratios using luminal A BC as a reference, does not allow us to estimate absolute BC subtype risk. So far, most population-based studies observed a significant inverse association between parity and hormone receptor-positive BC, although none of these studies took age into account [2, 17]. One recent study investigated whether associations between reproductive factors and premenopausal BC differed before and after age 40 years. The inverse association between parity and hormone receptor-positive BC was only observed in women aged > 40 years [27]. Most studies have not identified a statistically significant relationship between parity and the risk of TNBC [2], although one study in women aged between 20 and 44 years observed that a short time interval between menarche and FFTP was associated with an increased risk of TNBC [18]. Again, it should be noted that these studies did not assess differential effects by age or were focused on specific ages of diagnosis, and were often also hampered by the low prevalence of TNBC compared to hormone receptor-positive BC. Possibly, our findings may be explained by Russo’s hypothesis, which suggests that some BCs develop at a younger age (i.e., prior to the pregnancy) [4, 5]. Protection against hormone receptor-positive BC due to a pregnancy may thus result in a relative increase in TNBC at a later age, especially around 55 years, as suggested in this study.

The strength of this study is its large sample size with comprehensive data on molecular markers and other

| Table 4 Associations between age at menarche, age at FFTP and breast cancer subtypes |
|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|
|                 | Luminal B-like Odds ratio (95% CI) | P value | Luminal HER2-like Odds ratio (95% CI) | P value | HER2-like Odds ratio (95% CI) | P value | TNBC Odds ratio (95% CI) | P value |
| Nulliparous women | Age at menarche | Linear model | 1.06 (0.65–1.72) | 0.82 | 1.03 (0.59–1.78) | 0.92 | 1.19 (0.60–2.36) | 0.62 | 1.03 (0.59–1.80) | 0.92 |
|                  | +5 years | Linear model | 1.06 (0.65–1.72) | 0.82 | 1.03 (0.59–1.78) | 0.92 | 1.19 (0.60–2.36) | 0.62 | 1.03 (0.59–1.80) | 0.92 |
| Parous women     | Age at menarche | Linear model | 1.04 (0.84–1.28) | 0.74 | 0.99 (0.78–1.25) | 0.92 | 1.16 (0.87–1.56) | 0.31 | 0.99 (0.80–1.23) | 0.95 |
|                  | Menarche (+5 years) | Linear model | 1.04 (0.84–1.28) | 0.74 | 0.99 (0.78–1.25) | 0.92 | 1.16 (0.87–1.56) | 0.31 | 0.99 (0.80–1.23) | 0.95 |
|                  | Age at FFTP | Linear model | 0.78 (0.70–0.88) | <0.0001 | 0.93 (0.86–1.01) | 0.07 | 1.11 (0.96–1.28) | 0.15 |
|                  | (25 versus 20 years) | Linear model | 1.08 (1.00–1.16) | 0.049 | 1.01 (0.93–1.10) | 0.85 | 1.06 (0.95–1.17) | 0.28 | 0.93 (0.86–1.01) | 0.07 |
|                  | (30 versus 25 years) | Linear model | 1.08 (1.00–1.16) | 0.049 | 1.01 (0.93–1.10) | 0.85 | 1.06 (0.95–1.17) | 0.28 | 0.93 (0.86–1.01) | 0.07 |
|                  | (35 versus 30 years) | Linear model | 1.08 (1.00–1.16) | 0.049 | 1.01 (0.93–1.10) | 0.85 | 1.06 (0.95–1.17) | 0.28 | 0.93 (0.86–1.01) | 0.07 |
| Joint analysis age at menarche and age at FFTP | Age at menarche | Linear model | 1.01 (0.79–1.29) | 0.92 | 0.84 (0.64–1.09) | 0.19 | 1.06 (0.76–1.49) | 0.71 | 1.02 (0.80–1.29) | 0.90 |
|                  | Menarche (+5 years) | Linear model | 1.01 (0.79–1.29) | 0.92 | 0.84 (0.64–1.09) | 0.19 | 1.06 (0.76–1.49) | 0.71 | 1.02 (0.80–1.29) | 0.90 |
|                  | Age at FFTP | Linear model | 0.78 (0.69–0.88) | <0.0001 | 0.92 (0.85–1.00) | 0.045 | 1.09 (0.94–1.26) | 0.27 |
|                  | (25 versus 20 years) | Linear model | 1.07 (0.99–1.16) | 0.10 | 0.99 (0.91–1.09) | 0.91 | 1.03 (0.92–1.15) | 0.62 | 0.92 (0.85–1.00) | 0.045 |
|                  | (30 versus 25 years) | Linear model | 1.07 (0.99–1.16) | 0.10 | 0.99 (0.91–1.09) | 0.91 | 1.03 (0.92–1.15) | 0.62 | 0.92 (0.85–1.00) | 0.045 |
|                  | (35 versus 30 years) | Linear model | 1.07 (0.99–1.16) | 0.10 | 0.99 (0.91–1.09) | 0.91 | 1.03 (0.92–1.15) | 0.62 | 0.92 (0.85–1.00) | 0.045 |

The results of logistic regression models are reported for each breast cancer (BC) subtype while considering luminal A BC as a reference (the binary response takes values 0 for luminal A or 1 for the subtype that is considered), age at menarche or age at first full-term pregnancy (FFTP) are considered explanatory variables. A random intercept was introduced to account for clustering by study. All analyses reported in this section were performed with correction for age at diagnosis as a continuous variable. Age at menarche or age at FFTP were modeled linearly (for age at menarche), whereas linear and quadratic functions were considered for age at FFTP. The best fit was tested (as described in "Methods") and are reported here.

*HER2 human epidermal growth factor receptor-2, TNBC triple negative breast cancer, CI confidence interval, FFTP first full-term pregnancy

*Adjusted for age at diagnosis
detailed data derived from pathology reports. As such, we were able to derive BC molecular subtypes, which are known to differ in their prognosis, for all included patients. Importantly, we also used clinical criteria based on the 2011 St. Gallen report to refine the definition of the molecular subtypes. We failed, however, to detect obvious differences between luminal B and luminal HER2+ BC with the reproductive factors under study, suggesting that both groups in fact behaved similarly. Our sample size was slightly smaller compared to the previous BCAC study by Yang et al. [3], due to the fact that information on grade had to be available in addition to ER, PR and HER2 status to define subtypes according to the 12th St. Gallen International Breast Cancer Conference Expert Panel [22]. A potential limitation of this study is that data were derived from studies with various designs and methods to obtain risk factor and marker data. Furthermore, a relatively large proportion of patients were diagnosed at a young age, because some participating studies oversampled younger patients with BC and patients with a familial history of BC. In the future, since several new studies have joined BCAC and are in the process of providing more detailed reproductive risk factor data, we plan to take the effect of other variables such as body mass index, age at last pregnancy and breastfeeding into account. However, as we conducted case-case analyses by including a random effect for study in all models and also adjusted for age at diagnosis whenever applicable, it is unlikely that this may have affected our results.

Conclusion

We report that parity is differentially associated with BC subtypes and that the association for HER2-positive BC (relative to luminal A BC) depends on the patient’s age, but not on ER/PR status. Later age at FFPT was also inversely associated with TNBC, which suggests that an early pregnancy may increase the likelihood of developing TNBC relative to luminal A BC. Our results have to be confirmed in large population-based studies. However, they provide further support for different etiology between BC subtypes and suggest that models used to predict BC risk should take this into account.

Additional file

Additional file 1: Table S1, Definitions of breast cancer subtypes that have been applied in previous BCAC manuscripts. Table S2 Number of breast cancer patients with reproductive risk factor data in the 34 BCAC studies assessed in this study. Table S3 Number of breast cancer case patients with tumor marker data in the 34 BCAC studies assessed in this study. Table S4 Distribution of tumor characteristics according to breast cancer subtypes. Table S5 Association between parity (ever versus never) and BC subtypes for age overall and for specific ages (40, 50 and 60 years). Table S6 Frequency table showing parity by subtype and age group. Table S7 Associations between age at menarche, age at FFPT and breast cancer subtypes. The same analysis as in Table 4 is performed but here parity is considered a continuous variable. Table S8 Effect of parity (ever versus never) on BC subtype risk across all ages at BC diagnosis and corrected for BMI. Associations between age at menarche, age at FFPT and breast cancer subtype risk. (DOCX 60 kb)

Abbreviations

AIC: Akaike information criterion; BC: Breast cancer; BCAC: Breast Cancer Association Consortium; BMI: Body mass index; ER: Estrogen receptor; FFPT: First full-time pregnancy; HER2: Human epidermal growth factor receptor-2; PR: Progesterone receptor; TNBC: Triple negative breast cancer

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Authors' contributions
OL wrote the manuscript and made substantial contributions towards the conception and design of the analyses and the interpretation of data. AR wrote the manuscript and made substantial contributions towards the conception and design of the analyses and the interpretation of data. AL conducted the statistical analyses. RK made substantial contribution to the acquisition of data. MB made substantial contribution to coordinating the acquisition of data. QW made substantial contribution to coordinating the acquisition of data. AS made substantial contribution to the acquisition of data and interpretation of data, HW made substantial contribution to the acquisition of data and interpretation of data. ILA contributed to data acquisition. VA contributed to data acquisition. MWB contributed to data acquisition. JB contributed to data acquisition. CB contributed to data acquisition. SEB contributed to data acquisition. HB contributed to data acquisition. PB contributed to data acquisition. GB contributed to data acquisition. PH contributed to data acquisition. GCT contributed to data acquisition. JC contributed to data acquisition. SC contributed to data acquisition. FC contributed to data acquisition. AC contributed to data acquisition. SSC contributed to data acquisition. KC contributed to data acquisition. ME contributed to data acquisition. PAF contributed to data acquisition. JF contributed to data acquisition. HF contributed to data acquisition. GGG contributed to data acquisition. AGN contributed to data acquisition. PG contributed to data acquisition. PH contributed to data acquisition. AH contributed to data acquisition. JH contributed to data acquisition. HI contributed to data acquisition. MI contributed to data acquisition. DK contributed to data acquisition. JAK contributed to data acquisition. VMK contributed to data acquisition. JLI contributed to data acquisition. AL contributed to data acquisition. JEU contributed to data acquisition. AM contributed to data acquisition. SMM contributed to data acquisition. SMM contributed to data acquisition. KeMa contributed to data acquisition. KeMu contributed to data acquisition. HN contributed to data acquisition.
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References
1. DeSantis CE, et al. Cancer treatment and survivorship statistics, 2014. CA Cancer J Clin. 2014;64(4):252–71.
2. Anderson KN, Schwab RB, Martinez ME. Reproductive risk factors and breast cancer subtypes: a review of the literature. Breast Cancer Res Treat. 2014;144(1):1–10.
3. Yang XR, et al. Associations of breast cancer risk factors with tumor subtypes: a pooled analysis from the Breast Cancer Association Consortium Studies. J Natl Cancer Inst. 2011;103(3):250–63.
4. Russo J, et al. Breast differentiation and its implication in cancer prevention. Breast Cancer Res. 2005;112 Pt 2:5931–65.
5. Russo J, et al. The protective role of pregnancy in breast cancer. Breast Cancer Res. 2005;7(3):131–42.
6. Ma H, et al. Reproductive factors and breast cancer risk according to joint estrogen and progesterone receptor status: a meta-analysis of epidemiological studies. Breast Cancer Res. 2006;8(4):R43.
7. Lambe M, et al. Parity, age at first and last birth, and risk of breast cancer: a population-based study in Sweden. Breast Cancer Res Treat. 1996;38(3):305–11.
8. Adami HO, et al. Age at first birth, parity and risk of breast cancer in a Swedish population. Br J Cancer. 1980(42(5):651–8.
9. MacMahon B, Cole P, Brown J. Etiology of human breast cancer: a review. J Natl Cancer Inst. 1973;50(1):121–42.
10. Schedin P. Pregnancy-associated breast cancer and metastasis. Nat Rev Cancer. 2006;6(4):281–91.
11. Pathak DR. Dual effect of first full term pregnancy on breast cancer risk: empirical evidence and postulated underlying biology. Cancer Causes Control. 2002;13(4):295–8.
12. Clavel-Chapelon F, Gerber M. Reproductive factors and breast cancer risk. Do they differ according to age at diagnosis? Breast Cancer Res Treat. 2002;72(2):107–15.
13. Apter D, Vihko R. Early menarche, a risk factor for breast cancer, indicates early onset of ovulatory cycles. J Clin Endocrinol Metab. 1983;57(1):82–6.
14. Pike MC, et al. Estrogens, progestogens, normal breast cell proliferation, and breast cancer risk. Epidemiol Rev. 1993;15(1):17–35.
15. Collaborative Group on Hormonal Factors in Breast, C. Menarche, menstruation, and breast cancer risk: individual participant meta-analysis, including 118 964 women with breast cancer from 117 epidemiological studies. Lancet Oncol. 2012;13(1):141–51.
16. Li CI, et al. Timing of menarche and first full-term birth in relation to breast cancer risk. Am J Epidemiol. 2008;167(2):235–9.
17. Horn J, et al. Reproductive history and the risk of molecular breast cancer subtypes in a prospective study of Norwegian women. Cancer Causes Control. 2014;25(7):881–9.
18. Li CI, et al. Reproductive factors and risk of estrogen receptor positive, triple-negative, and HER2-neu overexpressing breast cancer among women 20-44 years of age. Breast Cancer Res Treat. 2013;137(2):579–87.
19. Gaudet MM, et al. Risk factors by molecular subtypes of breast cancer across a population-based study of women 56 years or younger. Breast Cancer Res Treat. 2011;130(2):587–97.
20. Ma H, et al. Pregnancy-related factors and the risk of breast carcinoma in situ and invasive breast cancer among postmenopausal women in the California Teachers Study cohort. Breast Cancer Res. 2010;12(3):R35.
21. Lord SJ, et al. Breast cancer risk and hormone receptor status in older women by parity, age of first birth, and breastfeeding: a case-control study. Cancer Epidemiol Biomarkers Prev. 2008;17(7):1723–30.
22. Goldhirsch A, et al. Strategies for subtypes–dealing with the diversity of breast cancer: highlights of the St. Gallen International Expert Consensus on the Primary Therapy of Early Breast Cancer 2011. Ann Oncol. 2011;22(8):1736–47.
23. Aerts M. Topics in modelling of clustered data. Monographs on statistics and applied probability. Boca Raton: Chapman & Hall/CRC; 2002. p. 308.
24. Phipps AI, et al. Reproductive and hormonal risk factors for postmenopausal luminal, HER-2-overexpressing, and triple-negative breast cancer. Cancer. 2008;113(7):1521–6.
25. Chung S, et al. Association between chronological change of reproductive factors and breast cancer risk defined by hormone receptor status: results from the Seoul Breast Cancer Study. Breast Cancer Res Treat. 2013;140(3):557–65.
26. Ritte R, et al. Reproductive factors and risk of hormone receptor positive and negative breast cancer: a cohort study. BMC Cancer. 2013;13:584.
27. Warner ET, et al. Reproductive factors and risk of premenopausal breast cancer by age at diagnosis: are there differences before and after age 40? Breast Cancer Res Treat. 2013;142(1):165–75.
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