Care competencies in adolescents: Development of a new measure and relationships with well-being

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Abstract
Worldwide, we observe an increasing focus on fostering well-being in adolescents. This is reflected in growing research in the field of care. However, empirical research is lacking that focuses on the relationship between components of care: receiving care, self-care and extending care. This lack of research is mirrored in the lack of valid measures for assessing care competencies in educational contexts. The present research, therefore, has four goals: 1) to create a valid and reliable scale assessing levels of care competencies based on the multifaceted and multidimensional concept of care; 2) to explore the relationship between socio-demographic characteristics and care competencies in adolescents; 3) to examine the psychological outcomes associated with different levels of care competencies in adolescents; 4) to explore the interaction of the three care dimensions in predicting adolescents’ well-being. Based on data from 742 adolescents, a reliable and valid scale could be developed with six subscales: Receiving care, self-care and extending care, with each of them falling apart in care competencies and care failures. Gender, age and academic achievement were related to care competencies/failures. Findings support the link between adolescents’ well-being and care competencies/failures. These results have implications for promoting adolescents’ well-being through school-based care-cultivation programs.

Keywords Adolescents · Care competencies · Well-being · Scale development

Adolescence is a developmental period characterized by rapidly transformative biological, neurological, psychological and social changes (Layden and Shale 2012). These dynamic transitions may place them at risk of ill-being (Bakar and Sidek 2013). Although the majority of adolescents report a high and stable level of well-being (Birkeland et al. 2012), a minority of young people report a lower level of well-being and mirror a number of subjective mental health worries. Around the world, about 20% of adolescents experience a behavioral problem or mental disorder (Kessler et al. 2005). This might be an under-estimation since adolescent’s mental health problems are often not recognized, untreated and/or not reported in research (Sourander et al. 2004). Ill-being not only has damaging effects on individual adolescents in terms of social, intellectual and emotional development but also impedes the society in its social security, sustainable development, healthcare cost and productivity (WHO 2016). A call for accelerated actions to improve the well-being of adolescents is, therefore, globally put forward (WHO 2019).

The opportunity to close the youth well-being gap is, however, real. The field of care has made significant progress in understanding the role of care competencies in promoting individuals’ well-being (Mind and Life Institute 2014). Accordingly, school-based programs which directly focus on enhancing care competencies for adolescents should be conducted (Noddings 2004). In this context, measuring and analyzing the care competencies of adolescents and their relationships with well-being may be a prerequisite for developing evidence-based interventions for adolescents. However, while
there are good theoretical reasons to believe that adolescents’ well-being depends on the quality and quantity of care that they are capable to receive from others, develop for themselves and extend to those around them, the construct of care has not yet been examined empirically. A key reason for this is the lack of psychometrically sound and age-appropriate measures to study care competencies in adolescents. Based on a literature review of the available empirically validated measures, MLI (Mind and Life Institute 2014) concluded no scale is yet available to measure care competencies in the adolescent population. This suggests future research needs to develop a valid and reliable measure of care competencies in adolescents; and based on this, empirically examine the well-being outcomes associated with different level of adolescents’ care competencies.

What Is Care? Toward an Integrated Definition

“The concept ‘care’ is commonplace in lay and sociological discourses but is hardly uniformly defined” (Thomas 1993, p. 649). Examination of the concept of care results in the identification of an overlapping set of perspectives (Morse et al. 1990). Nevertheless, most authors agree on the nature of care as either a feeling state (emotion, affection, love), caring about someone, or an activity state (work, tasks, labors), caring for someone (Morse et al. 1990). As such, care can be construed as both an attitude and an action. Held (2006) puts this as follows: “All care involves attentiveness, sensitivity and respond to needs” (p. 39). And Gastmans (2006) confirms: “Good care demands more than just good intention; good care ... is a practice of combining activities, attitudes, and knowledge of the situation” (p.137). Care, therefore, can also be approached as a competency since it combines knowledge, skills, attitudes and behavior, making an individual able to perform a certain task at a given level (Blömeke et al. 2015). As such, care can be structured following its cognitive, affective and behavioral competency facets: (1) being aware of and apprehending an individual’s current situation or needs (cognitive); (2) responding sympathetically and empathetically to those needs (affective); and (3) acting to meet the needs of an individual (behavioral).

Next to being multi-faceted, care is also a multi-dimensional concept. Looking at the actor-dimension, care rather can be characterized as an aspect of social relations than as an individual’s trait or behavior (Held 2006). This explains why Noddings (1984) stresses that – in a reciprocal relationship between two human beings – care implies both a person one-caring and a person cared-for. Care cannot be sustained if a caring person fails being attentive and/or refuses to respond, or if the person being cared-for is unable or unwilling to receive the care of others (Noddings 1984). This multi-dimensional nature helps enriching the concept of care with a receiving care and extending care dimension. Next, besides receiving care and extending care, also self-care is getting accepted as an indispensable dimension. Self-care implies monitoring of one’s own well-being, growth, and competencies in order to competently and successfully engage in a caring relationship (Hamington and Sander-Staudt 2011). It is essential to reflect on our own needs and ensure that oneself is not subsumed in a caring relationship (Tronto 1993).

Taken together, the MLI (Mind and Life Institute 2014) reflects the relational nature of care through a framework of three interdependent dimensions of care: receiving care, self-care and extending care. Building on the multifaceted conception of care, ‘receiving care’ can be seen as the competencies to be aware of and receptive to (1) others’ noticing to your own situation/needs (cognitive); (2) others’ empathetic feelings to your own situation/needs (affective); (3) others’ actions to meet your own needs (behavioral). ‘Self-care’ is regarded as the competencies to (1) notice and understand the needs of yourself (cognitive); (2) empathetically react to and accept your own situation/needs (affective); (3) act to meet your own needs (behavioral). ‘Extending care’ can be described as the competencies to (1) notice and understand others’ situation/needs (cognitive); (2) empathetically react to and accept others’ situation/needs (affective); (3) act to meet others’ needs (behavioral). From this perspective, empathy – the capacity to place oneself in another’s position (Bellet and Maloney 1991) – can be seen as part of the cognitive and affective facets of extending care. Mindfulness can be regarded as a cognitive and affective practice of self-care by which an individual becomes aware of and accepts painful experiences in a balanced and non-judgmental way (Brown and Ryan 2003). Especially, compassion, the desire and will to care for and to help alleviate one’s suffering (Gilbert et al. 2017), can be regarded as one typical, critical manifestation of extending care. In general, the above working definitions suggest that care is connected to other constructs such as social support, empathy, altruism, prosocial behavior, mindfulness and compassion (Nguyen et al. 2019).

What Is Well-Being?

Despite the available body of research, there is no consensus on the definition of well-being. Two approaches can be adopted in defining this concept: hedonism versus eudaimonism. The hedonic approach is based on the notion that increased pleasure and a decrease in pain lead to higher well-being (Ryan and Deci 2001). Accordingly, subjective well-being is defined as “the pursuit of happiness and a pleasant life” (Chen et al. 2012; p.1034) or the total of positive affect, negative affect and life satisfaction (Grenville-Cleave 2012). Alternatively, the eudaimonic approach is based on the premise that people attain a higher well-being when they
experience the meaningfulness of life, challenges, and personal
growth (Ryan and Deci 2001). Well-being within this ap-
proach is defined as psychological well-being (Chen et al.
2012). Burn (2016) stated that “psychological well-being re-
sfers to inter- and intra-individual levels of positive functioning
that can include one’s relatedness with others and self-referent
attitudes that include one’s sense of mastery and personal

growth” (p.1). Taken together, well-being is about: (1) the
presence of positive emotions and moods (e.g., contentment,

happiness); (2) the absence of negative emotions (e.g., stress,
depression, anxiety); and (3) satisfaction with life, fulfillment
and positive functioning. Based on this comprehensive per-
spective, though correlated, well-being is not direct opposites
of ill-being, which is related to pervasive negative affects and
functioning (Ryff et al. 2006). Specifically, those with low
level of well-being would not be expected to show high level
of ill-being.

Care Competencies and Well-Being
in Adolescents

Previous cross-sectional as well as longitudinal studies affirm
the important role of care competencies for adolescents’ well-
being. Regarding the receiving care dimension, several studies
(e.g., Bal et al. 2003; Danielsen et al. 2009) revealed that
social support perceived and received from families, teachers
or classmates increased adolescents’ life satisfaction and pos-
itive affect and decreased their depressive symptomatology
and negative affect. With respect to the self-care dimension,
studies by Bluth and Blanton (2015) and Neff and McGehee
(2010) showed that mindfulness and self-compassion were
significantly related to all measured dimensions of well-
being (reduced negative affect, perceived stress, anxiety and
depression; increased positive affect and life satisfaction). Regard-
ing the extending care dimension, previous studies showed
that adolescents who performed prosocial and altruistic
acts experienced significantly higher well-being (Layous
et al. 2012; Pareek and Jain 2012).

As a trainable resource, care competencies may provide a
valuable tool for promoting overall well-being (Jazaieri et al.
2017). Several intervention studies showed how cognitive and
affective empathy training reduced bullying (e.g., Sakofsky
2009), cyberbullying (e.g., Schultz-Keumhozl et al. 2016)
and aggressive behavior of adolescents (e.g., Vaziri and Lotfi
Azimi 2012). Besides, stronger mindfulness, compassion and
self-compassion interventions resulted in lower negative af-
fact and higher happiness, as well as more positive function-
ing of adolescents (Bluth and Eisenlohr-Moul 2017; Edwards
et al. 2014; Ferrari et al. 2019). More significantly, school-
based care-cultivation programs confirmed the impact of a
‘pedagogy of care’ on adolescents’ well-being (Battistich
et al. 1997; Chang 2008).

Moreover, as mentioned, care dimensions are interdepen-
dent; thus, they might interact with each other to affect indi-
viduals’ well-being (Mind and Life Institute 2014). For exam-
ple, extending care may be related to higher well-being only
when it flows naturally out of the other components: receiving
care and developing self-care; if not, it may lead to compas-
sion fatigue or empathy distress (Mind and Life Institute
2014). Empirically, findings in the study by Cosley et al.
(2010) revealed that the combination of received social sup-
port and compassion for others predicted decreases in stress
level. A cross-cultural multi-study by Hermanto et al. (2016)
showed that the ability to receive compassion from others
buffers the depressogenic effect of self-criticism.

In general, the number of empirically validated studies re-
lated to the relationship between care competencies and well-
being in adolescents tends to increase. However, to our knowl-
edge, no empirical studies have directly and simultaneously
examined the relationship between all three dimensions of
care competencies and well-being as well as their interactions
in predicting well-being in adolescents.

The Present Study

Based on the conceptualization of care competencies and re-
search evidence cited above, the present research, therefore
aimed to create a valid and reliable scale reflecting the multi-
dimensional nature of care and assessing these care-compet-
tencies. Further, we wanted to empirically test the following
hypotheses: (1) socio-demographic characteristics will be sig-
ificantly related to care competencies in adolescents; (2) care
competencies will be positively correlated with indicators of
well-being and negatively correlated with indicators of ill-
being in adolescents; (3) the three care dimensions will signif-
icantly interact with each other in predicting well-being in
adolescents.

Methods

Participants and Procedure

Adolescents, aged 13 to 18 years were recruited from two
junior high schools and two high schools in Hue City,
Vietnam. Active informed consent for school participation
was obtained through the directorate of the Education and
Training Department of Hue City in line with formal regula-
tions. A set of questionnaires to assess care competencies and
well-being was administered to 813 students. On the first page
of the questionnaire participants signed an active informed
consent. Administration took place during school time in reg-
ular class groups of 40 to 45 students, taking a maximum
45 min. Prior to data collection, survey proctors and teachers
were provided with the aims of the survey, an introduction to the concept of care and the instructions for administration. Next, teachers were responsible for the administration in their own class with the support of survey proctors. The first author of this article was available to answer questions raised by participants.

Prior to data analysis, a quality screening was set up. In cases of conflicting responses, too many missing items or strange answer patterns, tests were excluded. This resulted in 742 valid responses from 416 female and 326 male adolescents (M age = 14.63; SD = 1.08) to be included in the statistical analysis (91.3%). Demographic data of the participants are summarized in Table 1. Within this sample, only 0.3% of the data were missing. Little’s (1988) test, which tests the assumption of Missing Completely at Random (MCAR) versus Missing at Random (MAR) revealed that data were MCAR (normed χ² = 1.07). Therefore, in all further analyses we relied on the Full Information Maximum Likelihood approach (Enders 2001) and N = 742.

**Measures**

**Care Competencies** To measure adolescents’ care competencies, 36 potential scale items were first generated by the researchers, building on a literature review (Nguyen et al. 2019) and through brainstorming. In line with the three theoretical dimensions, three subscales were put forward: receiving care, self-care and extending care. Each subscale was composed of 12 items with 6 positively worded statements and 6 negatively worded statements, tapping into the three facets of care (cognitive, affective and behavioral facets). The use of an equal number of negative and positive worded items is recommended by several researchers to make measures more valid (e.g., Baumgartner and Steenkamp 2001; Podsakoff et al. 2003). Sample items are presented in Table 3. Respondents scored each item on a Likert scale from 1 totally untrue about me to 6 totally true of me. The validity and reliability of this Care Competencies Questionnaire for Adolescents (CCQA) is being presented in the first part of the Results section.

**Well-Being** Well-being of adolescents was measured using three validated scales. First, the self-rated version of the Strengths and Difficulties Questionnaire (SDQ- Goodman 1997) was used to assess emotional and behavioral disorders of adolescents. The SDQ is predictive of psychiatric diagnoses in many developed and developing countries, including Vietnam (Tran 2006). It is composed of 25 items, of which 20 items were used in this study, covering four areas of adolescents’ difficulties: emotional symptoms (5 items, e.g., Often unhappy, depressed or tearful), conduct problems (5 items; e.g., Often fights with other children), hyperactivity/inattention (5 items; e.g. Thinks things out before acting), and peer problems (5 items; e.g., Gets along better with adults). These items were rated on a 3-point Likert scale (0 = not true; 1 = somewhat true; 2 = certainly true). In line with the manual, a total difficulty score was calculated based on all 20 items (Goodman and Goodman 2012). Cronbach’s alpha in our data was .72. Second, the Perceived Stress Scale (PSS; Cohen and Williamson 1988) was used to measure adolescents’ general level of stress during the past month (e.g., In the past month, how often have you been upset because of something that happened unexpectedly?). Answers were rated on a Likert-scale ranging from 0 never to 4 very often. In Vietnam, this scale was underscored by Nguyen and Nguyen (2012) with satisfactory reliability (alpha = .71). In the present study, to obtain acceptable Cronbach’s alpha, we had to leave out four positive items with low item-total correlations.
Finally, six negative worded items were retained with a Cronbach’s alpha of .64. Third, the Satisfaction with Life Scale (SWLS, Diener et al. 1985) was used to assess overall life satisfaction of adolescents. The five scale items are phrased positively (e.g., In most ways my life is close to ideal) and were answered using a 7-point Likert-rating (1 strongly disagree to 7 strongly agree). In Vietnam, this scale was validated in medical university students by Nguyen and Nguyen (2012) who reported satisfactory reliability (alpha = .75). In the present study, Cronbach’s alpha was found to be acceptable (.66).

Data Analysis

In order to test the construct validity of the CCQA, an initial exploratory factor analysis (EFA) was performed with principal components extraction and promax rotation. Based on the EFA-findings, the factor structure was checked via confirmatory factor analysis (CFA), using robust maximum likelihood estimation (MLR) given non-normality of most item scores. Items not fitting the measurement model due to low factor loadings (< .30) were removed from the model, but keeping in mind the three facets of care (cognitive, affective and behavioral) to be covered in each subscale. The following fit indices were used: normed χ² (χ²/df), comparative fit index (CFI), standardized root-mean-square residual (SRMR), and root mean square error of approximation (RMSEA). The following benchmark fit values were applied: normed χ² ≤ 3, RMSEA < .05 (< .08 is acceptable), CFI > .95 (> .90 is acceptable), and SRMR < .05 (< .08 is acceptable) (Hu and Bentler 1999; Kline 2011). In order to test convergent validity, Pearson correlations were calculated between the factors.

Reliability of the scale scores was measured by calculating the internal consistency coefficient alpha (Cronbach 1951). Going beyond Cronbach’s alpha as a single index of reliability, we also evaluated average inter-item correlations as recommended by Clark and Watson (1995), which should fall between .15 and .50, as anything below .15 suggests being too broad of a construct while anything above .50 suggests redundancy.

In the second phase, inferential statistical analyses were conducted. First, multivariate analyses of variance (MANOVA) were used to test the relationships of care competencies with the sociodemographic variables. Considering the large sample size, a significance level of p < .01 was put forward (Kim 2015). Next, partial correlations and multiple linear regressions were applied to explore the relationships between care competencies and well-being, controlling for relevant demographic variables. Finally, two-way interaction effects were estimated using multiple regressions, to test the interaction of care dimensions in predicting adolescents’ well-being. Mplus version 8.3 (Muthén and Muthén 1998–2017) was used for CFA and SPSS 24.0 was used for the descriptive and inferential analysis.

Results

Factor Analysis of the CCQA

KMO (0.79) and Bartlett’s test (df = 630; χ² = 4156.10; p < .001) indicated the sample was adequate and factor analysis was valid. The EFA was separately conducted for each dimension of care. Analysis of eigenvalues, scree plot and interpretability of the factors suggested a two-factor solution for each of the three care competencies. After rotation, all the positively-worded items loaded on the first factor and all the negatively-worded items on the second factor, identifying two latent constructs underlying each dimension of care: care competencies and care failures. The two retained factors explained 36.6% of the variance of receiving care items, 31.8% of the variance of self-care items and 33.8% of the variance of extending care items.

Based on the two first-order factors (care competencies and care failures) resulting from EFA, a CFA with MLR estimation was performed to evaluate the fit of a two-factor model for each dimension of care separately. After stepwise removal of two items with poor factor loadings in each dimension of care (< .30), the model fitted the data sufficiently for self-care: normed χ² = 2.12; CFI = .91; RMSEA = .04; SRMR = .04. However, for receiving care and extending care scale, the model still did not fit the data adequately. Based on modification indices, the fit of these models was improved by allowing error correlations between items 3 and 8 (r = .38) and 5 and 9 (r = .23) of receiving care and items 25 and 26 of extending care (r = .48), resulting in good fit for the models for receiving care: normed χ² = 2.97, CFI = .93; RMSEA = .05; SRMR = .05, and extending care: normed χ² = 2.03; CFI = .95; RMSEA = .04; SRMR = .04. Next, an overall 6-factor CFA model was conducted on 30 items of the CCQA, based on the 2-factor models above. In this step, three additional error correlations (items 7 & 4, r = − .18, item 13 & 14, r = .19 and item 19 & 20 r = .22) were allowed based on modification indices for similarly worded items. As indicated in the top row of Table 2, this 6-factor model demonstrated good fit. Noticeably, as mentioned above, we decided to retain item 15 and 20 (< .30) to guarantee the measurement of the behavioral facet of self-care.

Some alternative models were also evaluated via CFA. First, a higher-order latent variable model was studied to determine if relationships between the six factors could be explained by two overarching constructs: care competencies and care failures. This hierarchical model resulted in unsatisfactory fit compared to the previous model (Δχ² = 148.91, p < .001; ΔCFI = .06). Second, based on the EFA-results showing a clear 2-factor structure (care competencies and care failures) in all three dimensions of care, a global two-factor model was estimated. However, as reflected in Table 2, compared to the 6-factor model, the model also did not fit the data.
sufficiently well, $\Delta \chi^2 = 272.72, p < .001$; $\Delta$CFI = .13. Third, another global model with three factors based on the variables rooted in the Mind and Life Institute (2014) framework of the care concept (receiving care, self-care and extending care) was tested. Similarly, this model did not fit the data adequately, $\Delta \chi^2 = 347.73, p < .001$; $\Delta$CFI = .15. In sum, comparison of all tested CFA models results in the conclusion that distinguishing between six factors within the CCQA best represents the data structure. The standardized loadings of this model are displayed in Table 3. Unstandardized factor score from the 6-factor model were saved for further analyses.

Internal consistency, descriptives and correlations of the CCQA are reported in Table 4. Despite some of the alpha’s being rather low, the mean inter-item correlation for each subscale fell within the recommended range. Most of the correlations between factors were meaningful and quite strong, mirroring convergent validity of the scale. However, we also observe a significant positive relationship between self-care competencies and self-care failures.

### Care Competencies and Sociodemographic Variables

MANCOVA was applied with eight sociodemographic variables (i.e., age, gender, family structure, family size, rank in family, family income, location, and academic achievement) as independent variables, and the care dimensions as dependent variables. Significant multivariate effects based on Wilk’s lambda were found for only two independent variables (i.e., gender and academic achievement). There was a strong connection between care competencies/failures and gender ($F(6,707) = 8.93, p < .001, \eta^2 = .07$), with higher receiving care competencies ($F(1,712) = 19.85, p < .001, \eta^2 = .03, b = -.12$), self-care competencies ($F(1,712) = 34.57, p < .001, \eta^2 = .05, b = -.20$), self-care failures ($F(1,712) = 16.05, p < .001, \eta^2 = .02, b = -.21$) and extending care competencies ($F(1,712) = 28.94, p < .001, \eta^2 = .04, b = -.24$) in female compared to male adolescents, and with higher extending care failures in male adolescents ($F(1,712) = 8.68, p < .01, \eta^2 = .01, b = .07$).

In terms of academic achievement, there was a statistically significant difference in care competencies/failures based on academic achievement ($F(6,707) = 4.01, p < .001, \eta^2 = .03$), indicating that those with good GPA had higher levels of self-care competencies ($F(1,712) = 14.44, p < .001, \eta^2 = .02, b = .09$); extending care competencies ($F(1,712) = 8.72, p < .01, \eta^2 = .01, b = .09$) and lower levels of extending care failures ($F(1,712) = 8.50, p < .01, \eta^2 = .01, b = -.05$).

### Care Competencies and Well-Being of Adolescents

In order to check whether control variables should be included in the correlation and regression analysis, a MANCOVA was run with the same eight sociodemographic variables as independent variables, and well-being indicators as dependent variables. Significant multivariate effects were only found for age ($F(3,710) = 5.00, p < .01, \eta^2 = .02$) and family economic condition ($F(3,710) = 7.07, p < .001; \eta^2 = .03$). Univariate results indicated that older adolescents experienced more emotional and behavioral difficulties ($F(1,712) = 7.64, p < .01, \eta^2 = .01, b = .03$) and perceived stress ($F(1,712) = 11.63, p < .001, \eta^2 = .02, b = .07$), compared to younger ones, and adolescents from richer families had higher life satisfaction than those from more poor ones ($F(1,712) = 9.94, p < .01, \eta^2 = .01, b = .22$). Based on these results, we decided to control for age and family economic condition in subsequent correlation and hierarchical regression analyses.

Next, partial correlations revealed that all three dimensions of care competencies were positively related with life satisfaction. Receiving care and extending care competencies were significantly negatively correlated with behavioral and emotional difficulties. Unexpectedly, extending care competencies and self-care competencies were found to be positively but relatively weakly correlated with perceived stress. All three dimensions of care failures were negatively related with life satisfaction and were positively related with perceived stress and emotional and behavioral difficulties (Table 5).

Subsequent hierarchical linear regressions with age and family economic condition entered in Step 1 of the regression and care competencies/failures in Step 2, showed that care competencies/failures explained an additional 18% of variance in perceived stress ($F(6,727) = 26.85, p < .001$). As can be seen in Table 5, the single predictor of higher perceived stress was more self-care failure. Regarding behavioral and emotional difficulties, care competencies/failures explained an additional 24% of variance ($F(6,727) = 39.45, p < .001$). Receiving care competencies, self-care failures and extending care failures were significant predictors of behavioral and emotional difficulties, in the direction as expected. In terms of life satisfaction, care competencies/failures additionally explained 12% of variance ($F(6,725) = 16.64, p < .001$).
However, none of the care competencies/failures appeared to be a unique predictor of life satisfaction.

**Interactions between Care Competencies/Failures Predicting Well-Being**

Three significant interactions were found out of 45 interactions tested. We used the simple slope plots by Dawson (2014) to interpret interaction effects. First, an interaction between self-care competencies and receiving care competencies emerged, predicting perceived stress in adolescents ($\beta = -.04, p < .01$). Figure 1a shows that the positive relationship between self-care competencies and perceived stress was weaker if receiving care competencies were high. Second, the interaction between extending care competencies and receiving care competencies of adolescents was significant ($\beta = -.05, p < .01$). As displayed in Fig. 1b, adolescents with high extending care competencies but low receiving care competencies, experienced most stress. Moreover, only when receiving care competencies were low, high extending care competencies predicted more stress. Finally, the effect of receiving care competencies on perceived stress also depended on the level of extending care failures ($\beta = -.05, p < .01$). As displayed in Fig. 1c, only adolescents with high receiving care competencies and high extending care failures, experienced more perceived stress.
Discussion

Though care competencies are believed to increase well-being in adolescents (Mind and Life Institute 2014), empirical evidence is lacking due to unavailability of age-appropriate instruments. This inspired the development of the CCQA to assess different dimensions of care in adolescents. This helped studying the unique association between care competencies, sociodemographic characteristics, and well-being, as well as the interactions between care dimensions in predicting adolescents’ well-being.

Validation of the CCQA

The findings from the EFA relatively unexpectedly revealed that items in each dimension of care did not cluster into three competency factors (cognition, affect, and behavior) as expected from Blömeke’s competency model (Blömeke et al. 2015). Instead, the positively and negatively worded items fell apart in two factors underlying each dimension: care competencies and care failures. According to Neff (2003), it is not uncommon for negative and positive items in self-report measures to load on separate factors. Neff (2003) had the same experience when developing and validating The Self-Compassion Scale. Recently, the studies by Brenner et al. (2018) and Van der Gucht et al. (2017) confirm that the positive (i.e., self-kindness, common humanity, and mindfulness) and negative (i.e., self-judgment, isolation, and over-identification) components of self-compassion reflect two distinct constructs (self-kindness and self-coldness). Gilbert’s Social Mentality Theory (Gilbert 2005) also points out that compassionate ways of responding to suffering tap into parasympathetic nervous system activity, and uncompassionate ways of responding tap into the sympathetic nervous system, suggesting that care competencies and care failures should not be measured as a single factor. Care failures, therefore, may not be simply the absence of care competencies. Feeling competent in care and failing in care are two different experiences. For example, a person who does not feel happy when noticing others are happy (low extending care competency), does not necessarily feel envious when others get happiness in their life (high extending care failure); or a person who does not tend to easily change his personal plans because of others’ plans (low self-care failure) does not necessarily make time for his hobbies (high self-care competency). Besides, an individual can exhibit both care competencies and care failures. For example, a person could be touched when seeing miserable and starving people but also sometimes could feel envious when others get achievements and happiness in their life.

Table 4  Internal Consistencies, Descriptives and Correlations of the CCQA Subscales

| Items          | α     | Average r_{item, item} | M   | SD | 2.   | 3.   | 4.   | 5.   | 6.   |
|----------------|-------|------------------------|-----|----|------|------|------|------|------|
| Receiving care |       |                        |     |    |      |      |      |      |      |
| Competencies   | .50   | .17                    | 4.33| .74| -.39***| .77***| -.06 | .73***| -.56***|
| Failures       | .72   | .34                    | 2.81| 1.00| -.04 | .12** | -.03 | .41***|
| Self-care      |       |                        |     |    |      |      |      |      |      |
| Competencies   | .49   | .16                    | 4.10| .81| .27** | .86***| -.56***|
| Failures       | .56   | .20                    | 3.16| .91| -.12**| .32***|
| Extending care |       |                        |     |    |      |      |      |      |      |
| Competencies   | .64   | .23                    | 4.25| .77| -.56***|
| Failures       | .50   | .20                    | 2.49| .84| -.82***|

Table 5  Correlations and Standardized Regression Coefficients among Care Competencies and Failures and Adolescents’ Well-Being & Prosocial Behavior in Adolescents

| Care Competencies        | Receiving care | Self-care | Extending care | Receiving care | Self-care | Extending care |
|--------------------------|----------------|-----------|----------------|----------------|-----------|----------------|
| SDQ Total Difficulties   | r   | -.18*** | -.02 | -.11** | .42*** | .46*** | .31*** |
| β                        | -.21*| -.08   | .29 | .31*** | .42*** | .09*  | .30** |
| Perceived stress         | r   | .05 | .19*** | .11** | .41*** | .42*** | .09*  |
| β                        | .05 | .09 | -.10 | -.01 | .43*** | -.06  |       |
| Life satisfaction        | r   | .29*** | .21*** | .23*** | -.25*** | -.18*** | -.25*** |
| β                        | .04 | .09 | .22 | -.20 | -.11 | .12 |       |

***p < .001; **p < .01; *p < .05
As a consequence, CFA confirmed the 6-factor model, distinguishing a competency and failure dimension in each of the three care dimensions. Model comparisons further confirmed both the differentiation of receiving care, extending care and self-care, as well as competencies and failures in each of these dimensions. Reliability of each of these factors based on Cronbach’s alpha in some cases was unsatisfactory, but average inter-item correlations were all within the expected range, confirming different items within each subscale measured the same construct but were not redundant (Clark and Watson 1995). The six separate but correlated factors assessed distinct but related components of care. The relationships between the six factors were quite logic and as such confirmed convergent validity of the new instrument as well as the interdependence of the three care dimensions as defined by Mind and Life Institute (2014). The ability to receive care empowers and is empowered by extending care to others; extending care flows naturally out of the feeling of being worthy (self-care) and being cared for by others (receiving care); and self-care is cultivated in the other two dimensions: receiving care and extending care (Mind and Life Institute 2014). However, unexpectedly, there was a moderate positive relationship between self-care competencies and self-care failures. As mentioned above, self-care competencies and self-care failures are
two distinct systems; adolescents can both feel competent in self-care and experience failure in self-care. This paradoxical finding, in fact, seems to reveal an ambivalent struggle inside some adolescents between self-coldness and self-kindness. Some adolescents seem to try to practice self-kindness at moments they tend to criticize and disappoint themselves (Brenner et al. 2018).

**Care Competencies and Sociodemographic Variables**

Out of eight sociodemographic characteristics, only gender and academic achievement were related to dimensions of care competencies and failures. As reported, girls possess higher levels of receiving care, self-care and extending care competencies and lower extending care failures, compared to boys. These results are consistent with previous findings indicating women have higher levels of care-seeking and care-giving than men (Hermanto and Zuroff 2016; López et al. 2018). Differences in extending care, self-care and receiving care between boys and girls are probably in large part due to different socialization processes, even nowadays and sure in Vietnam (Nguyen 2000), the country in which this study took place. For example, girls are handed dolls and baby carriages to play with (suggesting nurturing and caring behaviors) and boys are handed superheroes and toy soldiers (suggesting fighting, and protecting behaviors). As empathy is seen as a more “feminine” trait that is more acceptable for women, women may have learned to both communicate and recognize it more easily (Hertenstein and Keltner 2011). Noticeably, female adolescents also scored significantly higher in both self-competencies and self-care failures. Gender differences in self-care are not consistent in previous studies. Meta-analytic research suggests that women have slightly lower levels of self-compassion (Yarnell et al. 2015). However, Ayala et al. (2018) found out that men were more likely to report poorer self-care but women still experienced significantly more stress than men. On the one hand, these findings confirm again that self-care competencies and care failures are two different experiences. On the other hand, it seems to imply that female adolescents find self-care more challenging, compared to men. Females also reflect more self-blame and self-criticism when coping with stressful events compared to men (Kelly et al. 2008). However, ironically, the fear that self-compassion is a weakness prevents men from getting benefit from developing this skill (Yarnell et al. 2015).

Results revealed that adolescents with higher academic achievement also report higher receiving care, self-care and extending care competencies and lower levels of extending care failures. These findings support results from previous studies that students with adaptive social and emotional competencies are most likely to excel in school (e.g., Collie et al. 2018; McCormick et al. 2015). In general, our findings seem to suggest that future care cultivation programs should consider gender and academic achievement differences to set different priorities to participants.

**Care Competencies and Well-Being of Adolescents**

As expected, findings confirm that care competencies and failures and well-being are significantly related. The findings show a basic contribution of care competencies to well-being and a potential prediction of care failures to ill-being. All three dimensions of care failures were moderately or strongly associated with the three well-being indicators. In addition, noticeably, self-care failures are the strongest unique predictor of perceived stress and social-emotional difficulties. In contrast, the predictive power of care competencies in this present study was weak. Only receiving care competencies significantly predicted less emotional and behavioral problems. At first hand, the unexpected positive relationship between self-care and extending care competencies and perceived stress seems to imply that high level of these two competencies is somewhat stressful. However, these relationships were very weak and did not stay when controlling for care failures in the regressions. Therefore, these findings do not support the paradox of giving care (Lu and Argyle 1992) and unexpected side-effects of mindful self-awareness (Shapiro 1992).

Firstly, part of these findings may support the results of the study by Brenner et al. (2018) on an adult sample that self-coldness was more strongly related to depressive symptoms than self-kindness. Secondly, these findings support the conclusion from Bal et al. (2003), Bluth and Blaton (2015), Bluth and Eisenlohr-Moul (2017), Edwards et al. (2014), Ferrari et al. (2019), Layous et al. (2012), and Pareek and Jain (2012) that the strongest determinants of life-satisfaction, positive mental health are practical social support, self-compassion and care-giving practices; and from Stafford et al. (2015), Brenner et al. (2018) and Beaumont et al. (2016) that poor received social support, self-coldness and poorer compassion for others might lead to ill-being.

On the one hand, these findings imply that opening up to and cherishing the fact that other people care about them, helps adolescents to feel connected, needed, and included in the lives of others. Showing kindness, support and empathy for others helps them to strengthen their relationships with others and build new ones. Overall, both receiving care and extending care competencies can satisfy the need for relatedness by boosting social connection. In turn, a feeling of social connection to others might boost their sense of well-being. Research has shown that people who feel more connected to others reflect lower anxiety and depression rates (Seppala et al. 2013). Moreover, both receiving care and extending care competencies might also satisfy the need for competence of adolescents. Appropriate care from others can increase self-efficacy - an assurance that a person can achieve his desired goal (Akbarbegloo et al. 2015; Wang et al. 2015). Similarly, helping others will improve how adolescents see
themselves and value their contribution to the lives of others. In other words, extending care may give adolescents a sense of purpose in life and a feeling of self-worth. Self-care may enhance adolescents’ well-being because it also helps the need for competence to be fulfilled. By truly accepting themselves, being kind to themselves and by stopping criticizing themselves, adolescents tend to value themselves, which increases their self-efficacy.

On the other hand, the findings of this study also imply that the reluctance of adolescents to receive support from others and give support to others might limit their self-efficacy, their belongingness to others and prevent them from trusting to be cooperative sources of help, which in turn, decreases their well-being. Similarly, self-criticism and maladaptive perfectionism may put adolescents under pressure, which in turn leads to lower self-worth. These failures in receiving care, extending care and self-care, therefore, affect their well-being. Noticeably, self-care failures are the most important predictor of stress in adolescents. More focus should be put on the overcoming of failures in self-care, especially self-judgment in future care training and counseling programs.

Interactions between Care Competencies/Failures in the Prediction of Well-Being

As stated above, it might be the case that each dimension of care does not uniquely predict well-being, but that this relationship depends on the effect of other dimensions (Mind and Life Institute 2014). However, in this study, not much evidence for interaction between care dimensions was found. Only three interactions underscored the assumption that well-being depends on the ability to balance the different competencies of care. For example, the imbalance between giving care and receiving care seems to lead to more perceived stress. Adolescents with high extending care competencies but low receiving care competencies experienced most perceived stress; in contrast, adolescents with high receiving care competencies experience lower stress, regardless of their competency to extend care. These findings provide limited support to the conclusion from Cosley et al. (2010) that our ability to receive social support may be strengthened by the compassion for others, which may help us to adapt effectively to stress.

Limitations and Directions for Future Research

Several limitations of this study need to be acknowledged. First, the present study was entirely based on self-reports by adolescents, which might have inflated some of the findings, partly due to social desirable answering. Future research should include various data sources like questionnaires, observations, and structured interviews to reduce the impact of these biases.

Secondly, these findings are limited by the use of a cross-sectional design. Though cross-sectional data do not prevent us from assessing associations among variables, they constrain our ability to draw strong conclusions and to make causal inferences with regard to the well-being of adolescents. Therefore, future research should implement experimental (intervention) and longitudinal designs.

Thirdly, although we selected valid questionnaires, the problem with unreliability of some scales may occur due to the fact that part of the data was gathered right after the biggest flood since 1999 in Vietnam. This particularly might have influenced answers on specific items regarding care competencies and well-being. Future research with the CCQA should re-evaluate reliability and validity in other samples, as well as with other types of reliability, such as test-retest reliability.

Next, since the study was limited to 14–17-year-old adolescents, caution is needed to build on these data when looking at youngsters in early adolescence (11–13 years). We look forward to involving a larger age-range in future research.

Lastly, we linked care competences to well being. This was done in view of a further validation of the care competence scale. As such, we did not develop an extensive model about the relationship between care and well-being. This is clearly an avenue for future longitudinal research that can start now because an adequate measure is available.

Conclusion

The results of this study suggest that the CCQA provides an appropriate and valid instrument for measuring care competencies in adolescents in educational contexts. Six factors of the care competencies scale were identified: competencies and failures in receiving, extending and self-care. In general, the findings support the assumption that adolescents’ well-being depends on their ability and failure to receive care from others, to extend care to others, and to develop self-care, and also, their ability to balance these care dimensions. The former implies that our study contributes in a theoretical way to the literature by developing an integrated conceptualization of care and by putting forward empirical evidence. Also, our empirical results underpin the reliability and validity of an instrument in line with this new conceptualization. Our findings contribute to our understanding of the influences of care competencies in adolescents; e.g., that overcoming care failures seems more important than building care competencies in adolescents. Lastly, we could demonstrate the interdependence of the three care dimensions and how they interact with well-being. These contributions help grounding the importance and content of future care-cultivation programs targeting adolescents.

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Compliance with Ethical Standards

Conflict of Interest No conflict of interest in whatever form exists related to this study, for none of the authors. Moreover this study was not submitted before nor is it under consideration for publication elsewhere.

Ethical Statement Ethical approval for this study was obtained from Institutional Ethical Committee of Hue University.

Informed Consent “Active informed consent was obtained (1) the director of the Education and Training Department of Hue City, (2) the concemed school authority in line with formal regulations, and (3) from the participating adolescents themselves. See the original used informed consent form for adolescents (in Vietnamese) on the supplementary document.

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