Home and Workplace Neighborhood Socioeconomic Status and Behavior-related Health: A Within-individual Analysis

Auriba Raza, PhD1, Martin Claeson, MSc1 · Linda Magnusson Hanson, MD, PhD1 · Hugo Westerlund, PhD1 · Marianna Virtanen, PhD2 · Jaana I. Halonen, PhD1,3

Published online: 13 February 2021
© The Author(s) 2021. Published by Oxford University Press on behalf of the Society of Behavioral Medicine.
This is an Open Access article distributed under the terms of the Creative Commons Attribution License (http://creativecommons.org/licenses/by/4.0/), which permits unrestricted reuse, distribution, and reproduction in any medium, provided the original work is properly cited.

Abstract

Background The influence of individual and home neighborhood socioeconomic status (SES) on health-related behaviors have been widely studied, but the majority of these studies have neglected the possible impact of the workplace neighborhood SES.

Objective To examine within-individual associations between home and workplace neighborhood SES and health-related behaviors in employed individuals.

Methods We used participants from the Swedish Longitudinal Occupational Survey of Health who responded to a minimum of two surveys between 2012 and 2018. Data included 12,932 individuals with a total of 35,332 observations. We used fixed-effects analysis with conditional logistic regression to examine within-individual associations of home, workplace, as well as time-weighted home and workplace neighborhood SES index, with self-reported obesity, physical activity, smoking, excessive alcohol consumption, sedentary lifestyle, and disturbed sleep.

Results After adjustment for covariates, participants were more likely to engage in risky alcohol consumption when they worked in a workplace that was located in the highest SES area compared to when they worked in a workplace that was located in the lowest SES area (adjusted odds ratios 1.98; 95% confidence interval: 1.12 to 3.49). There was an indication of an increased risk of obesity when individuals worked in the highest compared to the time when they worked in the lowest neighborhood SES area (1.71; 1.02–2.87). No associations were observed for the other outcomes.

Conclusion These within-individual comparisons suggest that workplace neighborhood SES might have a role in health-related behaviors, particularly alcohol consumption.

Keywords Longitudinal study · Home neighborhood · Work neighborhood · Health-related behaviors · Socioeconomic status

Introduction

Individuals’ behaviors such as physical inactivity, sedentary lifestyle, smoking, alcohol consumption, and high body mass index (BMI) are important determinants of health and have been associated with an increased risk of morbidity and cardiovascular as well as all-cause mortality [1–3]. As these health-related behaviors are modifiable, they are crucial in preventing chronic diseases [4]. While strong individual-level socioeconomic gradients exist in these health-related behaviors [5–8], they may also be influenced by environmental characteristics of neighborhoods where people live, work, or spend their leisure time. In neighborhood-level studies, for instance, living in a low socioeconomic status (SES) neighborhood has been associated with binge drinking, physical inactivity, obesity, and smoking [5, 9–12].
Hypothesized mechanisms underlying these associations include environmental and contextual factors which capture how communities and neighborhoods affect individuals’ health and health-related behaviors [13, 14]. Neighborhood socioeconomics, for instance education, unemployment, poverty, and homeownership are major contextual factors that might drive both social factors and built characteristics of a neighborhood [14]. Studies have reported that neighborhoods with low SES might have low social cohesion, low social control, and high crime rates [15, 16]. Area-level social capital has been associated with physical activity [17], high crime rates to increased smoking prevalence [18] and increased BMI through decreased physical activity and decreased perceived safety [19]. Individuals with high or medium SES were reported to have a greater probability of good sleep quality compared to individuals with a low SES [20]. Anxiety, depression, and health status are important additional determinants, which are more prevalent in lower SES were observed to be associated with poorer sleep quality [20]. In addition, high densities of alcohol and smoking outlets and fast-food facilities, low esthetic qualities, and absence of adequate facilities for physical activity [21, 22] might lead to poor health-related behaviors of individuals living in low SES neighborhoods. While the mechanisms between associations of workplace neighborhood SES and health are likely similar to home neighborhoods, a majority of the research has been around home neighborhood SES [5, 9–12, 23], and less is known about impact of additional contextual environments such as those around workplaces, on behavior-related health [24–26].

People spend a considerable amount of waking time in their workplace. Workplace neighborhoods may thus play a role in shaping health-related behaviors. Few studies have reported findings of associations between workplace neighborhood and health [24, 27]. One linked healthy food availability around workplace environments with lower BMI [24] and another reported a higher density of physical activity facilities in workplace neighborhoods than in home neighborhoods [27]. Furthermore, the potential synergistic effect of home and work neighborhood SES on health-related behaviors has not been explored in a longitudinal setting [24].

The majority of studies on neighborhood health-related behaviors, especially the few examining workplace neighborhoods, have been cross-sectional [10, 12, 23–28], limiting the ability to draw a causal inference as the chronological order of exposure and outcome cannot be determined. These methodological limitations can be mitigated by a longitudinal study design with repeated measurements [29]. Within-individual analyses can assess whether individuals had healthier behaviors when they were living or working in a higher SES neighborhood compared to another time when the same individuals were living or working in lower SES neighborhoods [29, 30]. The within-individual design adjusts for unmeasured time-invariant characteristics of the individuals [29], such as personality, as individuals are compared with themselves when living or working in neighborhoods with different SES. Thus, a within-individual estimation can provide insight into potential causal processes linking neighborhood SES and health-related behaviors [9, 29].

In this study, we used repeated measures from a population-based survey to examine longitudinal associations of home and workplace neighborhood SES separately and jointly with health-related behaviors in gainfully employed individuals. We hypothesize that individuals when living or working in a higher SES neighborhood would have a lower likelihood of poor health-related behaviors compared to times they were living or working in lower SES neighborhoods.

Materials and Methods

Study Population

We used data from Swedish Longitudinal Occupational Survey of Health (SLOSH), a national longitudinal survey of work-life participation, work environment, and health and wellbeing that started in 2006 [31]. SLOSH participants consist of the respondents of the Swedish Work Environment Surveys 2003–2011 who were initially sampled biennially from 2003 to 2011 from the Labor Force Surveys by Statistics Sweden [32]. The SLOSH cohort participants, 16–64 years of age at the time of first response, are followed-up biennially and are sent two versions of the questionnaire: (i) gainfully employed for at least 30% of full-time or (ii) employed less than 30% or unemployed. Seven biennial waves of SLOSH from 2006 to 2018 have been conducted.

One of the objectives of this study is to determine the association of work neighborhood SES on behavior-related health. Therefore, we only included participants who were gainfully employed for at least 30% of full-time at the time of the survey. Further inclusion criterion was a response to outcome-related questions in at least two of the four SLOSH waves between 2012 and 2018. We excluded participants with missing information on covariates or objectively measured exposure. Total number of gainfully employed individuals who responded to any SLOSH wave between 2012 and 2018 was 14,618. We had an analytical sample of 12,932 participants after applying our exclusion criteria. This analytical sample contains individuals eligible for the analyses. Among the included there were slightly more women
Home and Work Neighborhood Socioeconomic Status

Our neighborhood units consist of areas within a Euclidean distance of a 500 or 1,000 m radius around the home and workplace address. Address information was from the end of December on the preceding year of the included surveys that were conducted in the spring of 2012, 2014, 2016, and 2018. Concept of SES is multidimensional and single variables do not capture the complexity of the concept of SES (whether at an individual or group level) [33]. Therefore, we used a summary neighborhood SES score as the main indicator of the SES of the neighborhood. To create the summary score, we used three variables representing the dimensions of income and social status; mean household income, low education (percentage of adults above 18 years with only elementary school education), and unemployment (percentage of unemployment) within 500 and 1,000 m radius of home or workplace address. These variables have been predominantly used to characterize neighborhood SES in the previous literature [11, 23, 24, 34–36]. These three socioeconomic determinants were obtained from Statistics Sweden. Since SLOSH surveys were conducted in spring, the data for above-mentioned three variables were obtained for the years preceding the SLOSH surveys, that is, 2011, 2013, and 2015 for the SLOSH waves 2012, 2014, and 2016, respectively. For the 2018 SLOSH wave, the reference year was 2016 as data were not available for 2017.

Prior to calculation of summary scores for neighborhood SES, we log-transformed mean income and educational attainment for a more normalized distribution. Low education and unemployment rates were coded as additive inverse to get the lowest values for the highest SES. For each SES determinant, standard z-scores (mean = 0, SD = 1) were derived. We calculated neighborhood SES by taking the mean values across the three z-scores. Using the mean values neighborhood SES was divided into quartiles, the highest quartile indicating the highest neighborhood SES. Our neighborhood SES measure was operationalized as quartiles as the majority of the neighborhood and behavior-related health research used quartiles and thus this facilitates comparison with the previous research [5, 9–11, 21, 23, 26]. Furthermore, our research question compares individuals moving from low to high neighborhood SES thus use of categorized measure was reasonable.

Summary measures for each participant’s “home and work” neighborhood SES were created using a previously applied approach [24]. We had information on the number of hours participants spent at work and commuting to and from work per week and used this information to estimate the number of hours spent at home per week. We calculated time-weighted averages of the home and work neighborhood SES. Weights were proportional to the number of hours spent at each location during the week.

Behavior-related Health Outcomes

From survey responses, we included physical inactivity, sedentariness, obesity (based on BMI), smoking, problem drinking, and disturbed sleep as outcomes. All outcome variables were collapsed into dichotomous variables as not all of them could be used as continuous variables due to categorized nature of the responses. The choice of cut-offs was based either on the previous literature or decided by the authors after accessing the information provided by the respondents.

Physical inactivity of participants was evaluated by using a question “How much do you exercise” including walking and cycling to and from work” with response alternatives: never exercise, move very little or take occasional walks, exercise now and then, and exercise regularly. We categorized participants as physically inactive if they selected the first or second response alternative and active if they selected the third or the last alternative [37].

Obesity was based on BMI that was calculated as self-reported weight in kilograms (kg) divided by self-reported height in meters squared in each follow-up. Obesity was dichotomized as Obese (BMI ≥ 30) versus not obese (BMI < 30).

Smoking status of participants was collected using a question, Do you smoke? and categorized as smokers (daily or occasional smokers), and nonsmokers (never or former smokers).

Risky alcohol consumption was assessed using the modified Cut-Annoyed-Guilty-Eye (CAGE) [38] questionnaire. It includes the following questions for participants reporting any drinking: (i) Have you felt you should cut down on your drinking? (ii) Have people annoyed you by criticizing your drinking? (iii) Have you felt bad or guilty about your drinking? (iv) Have you had a drink first thing in the morning to steady your nerves or get rid of a hangover? Participants’ alcohol consumption was categorized as risky if they reported a minimum of two problem drinking behaviors [38].

Disturbed sleep was based on the Disturbed Sleep Index from the Karolinska Sleep Questionnaire [39]. Participants were asked if they had difficulties falling asleep, restless sleep, repeated nocturnal awakenings, and premature awakening. For each question, there were six response alternatives: never, rarely, few times per month, 1–2 times per week, 3–4 times per week, and 5 or more
times per week. We defined disturbed sleep as having one or more sleep problems 3–4 times a week or more [37, 40].

Sedentariness: In the 2014, 2016, and 2018 waves, two questions were used to assess sedentary behavior: (i) “on average, how many hours do you spend sitting in a weekday; (a) during working day, (b) during travel time to and from work, and (c) during spare time.” For each part, there were five response alternatives: 0–1, 2–3, 4–5 h, 6–7, and 8 h or more and (ii) “How often do you usually take shorter breaks to move when sitting for longer hours? (a) at work, and (b) in spare time” with response alternatives; never/almost never, quite seldom, now and then, quite often, and frequently. We only included participants who responded to all sections of the above two questions. The mean of the hours spent sitting during the working day, commuting, and spare time were in each wave used as cutoffs for that wave. We formed a combination variable based on the two questions (sitting hours and taking breaks). Participants’ behavior was characterized as sedentary if they spent 7.5 (mean) hours or more sitting without taking breaks quite often and active if otherwise. Previous studies on sedentary behavior categorized participants’ behavior as sedentary based on reading of accelerometers or inclinometers or daily diaries separately or combined [41] but our choice of categorization was derived by the information on sitting time provided by study participants.

Covariates

Self-reported sociodemographic variables were a number of children (one or more children under 12 years) and occupational position (low = manual employees, intermediate = nonmanual employees, high = professionals, and self-employed). This categorization was based on the Swedish socioeconomic classification [42] and study participants were categorized based on the information provided about occupation, detailed job title, and main tasks at work. Information on age and civil status (married/cohabiting vs. not) was obtained from registers.

Health and work-related variables included chronic diseases, symptoms of depression, and psychological job strain. For chronic diseases, if participants had hypertension or cardiovascular disease or diabetes or rheumatic disorders or musculoskeletal disorders during the past 2 years they were coded as having chronic disease [37].

In all waves, symptoms of depression were evaluated using a six-item subscale of the (Hopkins) Symptom Checklist (SCL) resulting in SCL-Core Depression scale [43, 44]. Respondents were asked to score on a five-category scale the extent that they felt blue, had no interests in things, were lethargic or low in energy, were worrying too much about things, blamed themselves for things, and felt everything is an effort. For each item, scores were summed to get a continuous scale assessing the severity of depression. Using the continuous scale, we created a binary variable with a cut-off score of ≥17 [44]. A score of 17 has been identified as the best cut-point for major depression (sensitivity 0.68, specificity 0.98) which predicted subsequent purchases of antidepressants as well as hospitalizations with a depressive episode [44].

Job strain is defined as high job demands and low job control and heavy drinkers are more likely to report job strain [45]. Questions assessing job strain were based on Karasek’s job demand-control model in all waves [46]. A Demand Control Questionnaire has five job demand items: (i) working fast, (ii) working hard/intensively, (iii) no excessive amount of work/too much effort, (iv) enough time, and (v) conflicting demands, while job control has six control items: (a) learn new things, (b) high level of skill, (c) creativity/initiative, (d) repetitive work, (e) a lot of say/what to do, and (f) little freedom/how to do. In each wave for each participant, we created a mean of all five job demand and six job control items and used a median of means as a cut-off point. Participants with mean demands scores above the median and mean control scores below the median were categorized as having job strain [47].

These covariates were chosen as they have been linked to neighborhood SES and behavior-related health [26, 34, 45, 46].

Statistical Analyses

We applied a fixed-effects approach [48], also known as within-individual or case-crossover design, and conditional logistics regression to analyze associations between home- and work-neighborhood SES and health-related behaviors in gainfully employed individuals. We used three exposure measures, that is, home and work neighborhood SES in separate models and as a time-weighted (“home and work”) variable. The main exposure variable used in the analyses was neighborhood SES within 500 m radius of home and workplace address. Small neighborhood of 500 m, on average 6–10 min walking distance from the addresses might correlate well with how residents perceive their neighborhoods.

We included participants who responded to a minimum of two waves as the within-individual method requires information from participants who report unfavorable health behavior (“case” situation) in one wave and favorable health behavior (“control” situation) in another wave. Those whose neighborhood SES changed between the waves were included in each analysis as “informative.” Since each case serves as its own control, time-invariant individual characteristics are controlled by the design [48].

In the analyses, our first model (Model 1) included adjustments for age, occupational position, marital status,
and presence of children as covariates. In subsequent models, we additionally adjusted for chronic disease and depressive symptoms (Model 2); and job strain (Model 3). All covariates were included as time-varying variables.

Sex differences have been reported in prevalence of sleep problems [49], alcohol consumption [50], and smoking [11]. Therefore, possible effect modification of the association between neighborhood SES and behavior-related health by sex was investigated by constructing multiplicative interaction terms.

As sensitivity analyses, for health-related behaviors demonstrating associations with the main exposure, we ran home neighborhood SES models adjusted for work neighborhood SES and vice versa. Data on socioeconomic determinants was based on income, education attainment, and employment status of individuals living in the neighborhoods. In 2010 in Sweden, about 12% of workplace areas with concentrations of workplaces with a minimum of 50 employees (e.g., industrial areas, mining areas, healthcare institutions, airports, nuclear power plants, and military installations) located outside urban areas [51]. Thus, not all workplace neighborhoods lie within residential areas, and we had missing data for the workplace neighborhood income, low education and unemployment. For sensitivity analyses, we imputed missing values with mean values from the year in question before creating the summary SES score to test if gaining statistical power changed our results. Finally, we ran analyses using a larger, 1,000 m buffer for the neighborhood exposures.

Lowest SES quartile was used as a reference. Effect estimates are expressed as odds ratios (OR) with 95% confidence interval (CI). All analyses were performed using SAS 9.4.

### Results

A total of 12,932 participants responded to a minimum of two surveys between 2012 and 2018 and had information...
on all covariates and exposure. Descriptive statistics of the study population are provided in Table 1. The participants had a mean age of 50 years (range 20–75) at the time of their first response and slightly more than half (57%) were women. Over 82% of the participants were working full time and a majority (48%) held intermediate occupational positions. Fourteen percent of the participants both lived and worked in a neighborhood with the highest SES, while 3.6% lived in the highest but worked in the lowest SES neighborhood. The unemployment rate was very low.
in both home and work neighborhoods (Table 1). There were 3,224 individuals who had a change in their home neighborhood SES between two survey waves: 2,731 had one, 464 had two, and 29 had three changes. Work neighborhood SES was changed among 3,684 individuals: 2,935 had one, 696 had two, and 53 had three changes.

We observed that the SES of the workplace neighborhood was associated with risky alcohol consumption (Fig. 1). When the participants worked in a high SES neighborhood they were more likely to engage in risky alcohol consumption compared to the time when their workplace was located in the low SES neighborhood (fully adjusted OR 1.98, 95% CI: 1.12, 3.5). Although the SES of the home neighborhood had a significant negative association with risky alcohol consumption in the second-lowest SES quartile, estimates approached one as the participants moved to the higher SES home neighborhood (Fig. 1). For the time-weighted home and work neighborhood SES, effect estimates indicated that higher neighborhood SES was associated with an increased risk of risky alcohol consumption, however, associations were not statistically significant (Fig. 1). The associations between work and home neighborhood SES and risky alcohol consumption were not modified by sex (p-values for interaction ranged between .3 and .9).

Regarding the other outcomes (Table 2), we observed an indication that participants might be at an increased risk of obesity when their workplace was located in the highest SES neighborhood compared to a time when their workplace was located in a neighborhood with the lowest SES (fully adjusted OR 1.71, 95% CI 1.02 to 2.87). However, no associations were observed between home and time-weighted average home and work neighborhood SES and obesity (Tables 3 and 4).

We observed no associations for the three neighborhood SES variables with physical inactivity, smoking, sedentariness, or disturbed sleep (Tables 2–4).

**Sensitivity Analyses**

When associations between home neighborhood SES and risky alcohol consumption were adjusted for work neighborhood SES, estimates remained similar but confidence intervals became wider as the number of observations

| Outcomes   | Work neighborhood socioeconomic status | First (ref) quartile | Second quartile OR (95% CI) | Third quartile OR (95% CI) | Fourth quartile OR (95% CI) |
|------------|----------------------------------------|----------------------|-----------------------------|-----------------------------|----------------------------|
| Obesity (2,646) |                                        |                      |                             |                             |                            |
| Model 1a   | 1                                      | 1.33 (0.94–1.88)     | 1.04 (0.68–1.59)            | 1.65 (0.99–2.77)            |
| Model 2b   | 1                                      | 1.31 (0.92–1.86)     | 1.02 (0.67–1.56)            | 1.64 (0.98–2.74)            |
| Model 3c   | 1                                      | 1.33 (0.93–1.89)     | 1.05 (0.68–1.61)            | 1.71 (1.02–2.87)            |
| Physical inactivity (5,374) |                                  |                      |                             |                             |                            |
| Model 1a   | 1                                      | 1.04 (0.83–1.31)     | 1.04 (0.80–1.36)            | 1.09 (0.80–1.48)            |
| Model 2b   | 1                                      | 1.04 (0.82–1.30)     | 1.04 (0.80–1.36)            | 1.09 (0.80–1.48)            |
| Model 3c   | 1                                      | 1.03 (0.82–1.30)     | 1.04 (0.80–1.36)            | 1.09 (0.80–1.48)            |
| Smoking (1,362) |                                  |                      |                             |                             |                            |
| Model 1a   | 1                                      | 1.20 (0.72–2.00)     | 0.63 (0.33–1.19)            | 1.28 (0.62–2.65)            |
| Model 2b   | 1                                      | 1.20 (0.72–2.00)     | 0.63 (0.33–1.21)            | 1.29 (0.62–2.68)            |
| Model 3c   | 1                                      | 1.20 (0.72–2.01)     | 0.63 (0.33–1.21)            | 1.29 (0.62–2.68)            |
| Sedentariness (4,694) |                                |                      |                             |                             |                            |
| Model 1a   | 1                                      | 0.91 (0.71–1.17)     | 0.84 (0.62–1.15)            | 0.79 (0.55–1.12)            |
| Model 2b   | 1                                      | 0.90 (0.70–1.16)     | 0.85 (0.62–1.17)            | 0.80 (0.56–1.14)            |
| Model 3c   | 1                                      | 0.91 (0.70–1.17)     | 0.86 (0.63–1.18)            | 0.81 (0.57–1.16)            |
| Disturbed sleep (8,120) |                              |                      |                             |                             |                            |
| Model 1a   | 1                                      | 0.94 (0.77–1.15)     | 0.96 (0.76–1.21)            | 1.02 (0.78–1.35)            |
| Model 2b   | 1                                      | 0.92 (0.76–1.13)     | 0.97 (0.76–1.23)            | 1.06 (0.80–1.40)            |
| Model 3c   | 1                                      | 0.92 (0.75–1.12)     | 0.97 (0.76–1.23)            | 1.06 (0.78–1.40)            |

*a*Adjusted for age, occupational position, marital status, and presence of children under 12.

*b*Adjusted for age, occupational position, marital status, presence of children under 12, chronic disease, and depressive symptoms.

*c*Adjusted for age, occupational position, marital status, presence of children under 12, chronic disease, depressive symptoms, and job strain.
dropped (3,038 vs. 2,096) due to missing data on work neighborhood SES (Supplement Table 1). However, associations between work neighborhood SES and risky alcohol consumption remained robust after adjusting for home neighborhood SES. A similar pattern was observed for work neighborhood SES and obesity. After imputing zero values with mean values in the workplace neighborhood SES determinants, the effect estimates for risky alcohol consumption were attenuated but remained significant for the highest SES category (Supplement Table 2). For obesity, associations in the highest SES quartile became weaker and nonsignificant.

Results for 1,000 m buffer demonstrated a higher likelihood of risky alcohol use when participants lived in a high SES neighborhood compared to the time when they lived in a low SES neighborhood, however, confidence intervals were very wide and nonsignificant (Supplement Table 3). For obesity, associations in the highest SES quartile became weaker and nonsignificant. For obesity, work neighborhood SES demonstrated similar patterns as the 500 m buffer but of a weaker magnitude (Supplement Table 3).

**Discussion**

In this within-individual study, we observed associations between SES of work neighborhood and an increased risk of risky alcohol consumption in gainfully employed individuals. For home neighborhood SES, the effect estimates were mainly in the opposite direction though nonsignificant for the highest SES quartiles. There was also an indication that while working in the highest SES neighborhoods individuals might more likely be obese compared to the time when they were working in the lowest SES neighborhood. No associations were observed for the other outcomes.

In our study, we observed no consistent associations between home neighborhood SES and health-related behaviors. Although living in the second lowest versus lowest home neighborhood SES quartile was associated with a decreased likelihood of risky alcohol consumption, the association was not consistent through the SES quartiles. Large systematic reviews comprising mainly of cross-sectional studies, have reported limited and conflicting support for the association between area-level disadvantage and increased alcohol use with some
findings supporting the hypothesis while others pointing in the opposite direction [28, 52, 53]. Longitudinal studies with varying methodological approaches have reported associations of increased neighborhood poverty and neighborhood SES with binge drinking [54] and decrease in weekly alcohol consumption [55], respectively. Conversely, high urbanization and high SES have also been associated with increased risky alcohol consumption [30]. The inconsistent home neighborhood SES results might stem from heterogeneity among studies in terms of analytical methods applied [29], classification of alcohol intake [56], differences in socioeconomic gradients and neighborhood structures in different countries, as well as differences in study populations [29, 56]. We used a within-individual approach, that is, fixed-effect method, which is effective in reducing the confounding effects of unmeasured time-invariant causes. However, its focus on within-individual variance reduces its statistical power [29] and might thus provide weaker effect estimates [30]. Moreover, it is likely that we might not have captured all hazardous drinkers since our study participants were gainfully employed which already puts them in a higher stratum in the socioeconomic gradient.

Interestingly, our results for work neighborhood SES and alcohol use and obesity were in a different direction than those reported in previous studies regarding home neighborhood SES. This was also contrary to our hypothesis, though the hypothesis was based mainly on studies using home neighborhood SES. To the best of our knowledge, this is the first within-individual study to investigate the relationship between work neighborhood SES and behavior-related health. Previously, only two studies have considered work neighborhood SES and both were cross-sectional. One study focused on school teachers in Finland and reported an increased risk of heavy alcohol consumption in teachers working in poor vs. wealthy neighborhoods [26]. Our study sample represents working Swedish population from all occupations and these workplaces might have more variability in work and social cultures. It could be that social comparisons at workplace might be underlying the observed associations of this study. As a post-hoc analysis, we checked if individual SES had an interaction with work neighborhood SES in the model for risky alcohol consumption. Such interaction did not exist (p-value = .93) and thus social comparison

### Table 4. Associations between weighted average of home and work neighborhood socioeconomic status and health-related behaviors among gainfully employed individuals

| Outcomes          | Home and work neighborhood socioeconomic status (quartile) | First (ref) quartile | Second quartile OR (95% CI) | Third quartile OR (95% CI) | Fourth quartile OR (95% CI) |
|-------------------|----------------------------------------------------------|----------------------|-----------------------------|-----------------------------|-----------------------------|
| **Obesity (2,487)** |                                                          |                      |                             |                             |                             |
| Model 1<sup>a</sup> | 1                                                        | 1.12 (0.70–1.80)     | 0.99 (0.54–1.79)            | 0.87 (0.42–1.79)            |
| Model 2<sup>b</sup> | 1                                                        | 1.14 (0.71–1.82)     | 0.97 (0.53–1.76)            | 0.85 (0.41–1.76)            |
| Model 3<sup>c</sup> | 1                                                        | 1.12 (0.70–1.80)     | 0.97 (0.53–1.76)            | 0.84 (0.41–1.74)            |
| **Physical inactivity (5307)** |                                                          |                      |                             |                             |                             |
| Model 1<sup>a</sup> | 1                                                        | 1.05 (0.80–1.39)     | 0.91 (0.64–1.28)            | 1.08 (0.70–1.65)            |
| Model 2<sup>b</sup> | 1                                                        | 1.04 (0.79–1.37)     | 0.90 (0.64–1.27)            | 1.07 (0.70–1.64)            |
| Model 3<sup>c</sup> | 1                                                        | 1.04 (0.79–1.38)     | 0.90 (0.64–1.27)            | 1.07 (0.70–1.64)            |
| **Smoking (1,264)** |                                                          |                      |                             |                             |                             |
| Model 1<sup>a</sup> | 1                                                        | 0.69 (0.32–1.47)     | 0.79 (0.29–2.12)            | 1.64 (0.52–5.18)            |
| Model 2<sup>b</sup> | 1                                                        | 0.69 (0.32–1.47)     | 0.79 (0.29–2.12)            | 1.62 (0.51–5.13)            |
| Model 3<sup>c</sup> | 1                                                        | 0.70 (0.32–1.48)     | 0.79 (0.29–2.13)            | 1.63 (0.52–5.14)            |
| **Sedentariness (4,370)** |                                                          |                      |                             |                             |                             |
| Model 1<sup>a</sup> | 1                                                        | 0.91 (0.62–1.34)     | 0.95 (0.59–1.55)            | 1.05 (0.59–1.86)            |
| Model 2<sup>b</sup> | 1                                                        | 0.89 (0.61–1.31)     | 0.95 (0.59–1.55)            | 1.03 (0.58–1.83)            |
| Model 3<sup>c</sup> | 1                                                        | 0.90 (0.61–1.32)     | 0.96 (0.59–1.55)            | 1.04 (0.59–1.85)            |
| **Disturbed sleep (7,591)** |                                                          |                      |                             |                             |                             |
| Model 1<sup>a</sup> | 1                                                        | 1.02 (0.78–1.33)     | 1.06 (0.76–1.48)            | 1.20 (0.73–1.65)            |
| Model 2<sup>b</sup> | 1                                                        | 1.01 (0.77–1.33)     | 1.07 (0.76–1.49)            | 1.09 (0.72–1.65)            |
| Model 3<sup>c</sup> | 1                                                        | 1.02 (0.78–1.34)     | 1.08 (0.77–1.51)            | 1.09 (0.72–1.66)            |

<sup>a</sup>Adjusted for age, occupational position, marital status, and presence of children under 12.

<sup>b</sup>Adjusted for age, occupational position, marital status, presence of children under 12, chronic disease, and depressive symptoms.

<sup>c</sup>Adjusted for age, occupational position, marital status, presence of children under 12, chronic disease, depressive symptoms, and work strain.
effects, that is, employees with lower individual SES would drink more if they see a richer environment and people during the day, is unlikely to explain our finding. However, likely explanation could be that in some occupations after-work drinks with colleagues to decompress a long day or a week, celebrating wins or new clients, welcomes and farewells to a colleague, and weekend parties at pubs may be more common and considered important to climb the corporate ladder. Availability of alcohol may be another factor as higher SES workplace neighborhoods are often concentrated with on-premise and off-premise alcohol outlets making access to alcohol for people working in these neighborhoods easier. Overall, it is likely that competitive, stressful or nonrewarding work in combination with social norms and values that support drinking [57] along with easy access to alcohol can lead to risky drinking. The second prior study investigated BMI in relation to workplace neighborhood SES in a working population of over 50 years of age and did not report significant associations [24]. In our study sample, including participants from all working-age groups, higher workplace neighborhood SES indicated some association with obesity. However, the association was not robust to different testing, so further studies are needed to confirm if such association exists. Importantly, future studies should consider associations between workplace neighborhood SES and health behaviors in urban areas only as we had no data to perform such restricted analyses.

This study has several strengths but also limitations. The major strength of our study is the within-individual study design with repeated measurements. This allowed us to examine if changes in neighborhood SES were associated with changes in health-related behaviors that strengthen causal inference. While the fixed effects’ focus on within-individual variation might cause a lack of precision, it is likely to provide less confounded estimates as it controls for many unmeasured time-invariant factors by design [29]. However, we cannot exclude potential confounding from unmeasured time-varying confounding. Another strength of our study is the use of a population-based sample. Previous studies that examined both home and non-home neighborhoods were generally confined to study populations from one geographical area whereas our study population is more geographically representative of the working Swedish population. This, however, limits the generalizability of our results regarding home neighborhoods to unemployed, student or retired populations. It is also likely that we did not capture all hazardous drinkers if they are more prevalent among the non-working populations.

We had data on the number of hours participants spent on workplace and commuting between home and workplace which enabled us to estimate time-weighted average of time spent at each neighborhood. We had some missing data for workplace neighborhoods that did not have residential buildings within 500 m radius around them. However, when missing data were imputed with mean values of the workplace neighborhood SES our results persisted. Moreover, for all neighborhood SES variables we used reliable routinely collected register-based statistics from Statistics Sweden. This information along with home and workplace addresses was available for end of each year. As the surveys were conducted in the spring we used exposure data from the end of the previous year. This mismatch and the lack of information on all address changes among the study participants might have caused some exposure misclassification. All our outcome measures were self-reported which might introduce social desirability bias [58], but the within-individual design reduces this type of bias as participants’ response styles are likely to be similarly biased from one survey to another.

Conclusions

Workplace neighborhood SES was associated with a higher risk of risky alcohol consumption. Findings of our study suggest that the consideration of work neighborhood SES along with home neighborhood SES might enhance our understanding of how environments where individuals spend majority of their waking time influence their health-related behaviors. Also, in order to get more reliable estimates regarding the associations between neighborhood characteristics and health, more longitudinal within-individual comparisons are needed.

Acknowledgments This study was supported by Swedish Research Council for Health, Working life and Welfare (project 2018-00479).

Compliance with Ethical Standards

Authors’ Statement of Conflict of Interest and Adherence to Ethical Standards The authors declare that they have no conflict of interests.

Authors’ Contributions A.R., M.C., L.M.H., H.W., M.V., and J.I.H. conceived and designed the experiments; A.R. conducted the statistical analyses and drafted the article; M.C., L.M.H., and H.W. contributed to the data and materials; L.M.H., H.W., and J.I.H. contributed to the funding of the study. All authors were involved in writing and approved the submitted version.

Ethics Approval and Consent to Participate The Ethical Review Board has approved this study (Dnr: 2019-01272).

References

1. Andersen LB, Schnohr P, Schroll M, Hein HO. All-cause mortality associated with physical activity during leisure time, work, sports, and cycling to work. Arch Intern Med. 2000;160:1621–1628.
2. Biswas A, Oh PI, Faulkner GE, et al. Sedentary time and its association with risk for disease incidence, mortality, and hospitalization in adults: A systematic review and meta-analysis. *Ann Intern Med*. 2015;162:123–132.

3. Kvaavik E, Batty GD, Ursin G, Huxley R, Gale CR. Influence of individual and combined health behaviors on total and cause-specific mortality in men and women: The United Kingdom health and lifestyle survey. *Arch Intern Med*. 2010;170:711–718.

4. Lloyd-Jones DM, Hong Y, Labarthe D, et al.; American Heart Association Strategic Planning Task Force and Statistics Committee. Defining and setting national goals for cardiovascular health promotion and disease reduction: The American Heart Association’s strategic Impact Goal through 2020 and beyond. *Circulation*. 2010;121:586–613.

5. Giskes K, Turrell G, Bentley R, Kavanagh A. Individual and household-level socioeconomic position is associated with harmful alcohol consumption behaviours among adults. *Aust N Z J Public Health*. 2011;35:270–277.

6. Pampel FC, Krueger PM, Denney JT. Socioeconomic disparities in health behaviors. *Annu Rev Sociol*. 2010;36:339–347.

7. Stringhini S, Dugravot A, Shipley M, et al. Health behaviours, socioeconomic status, and mortality: Further analyses of the British Whitehall II and the French GAZEL prospective cohorts. *PLoS Med*. 2011;8:e1000419.

8. Stringhini S, Sabha S, Shipley M, et al. Association of socioeconomic position with health behaviors and mortality. *JAMA*. 2010;303:1159–1166.

9. Boone-Heinonen J, Diez Roux AV, Kiefe CI, Lewis CE, Guilkey DK, Gordon-Larsen P. Neighborhood socioeconomic status predictors of physical activity through young to middle adulthood: The CARDIA study. *Soc Sci Med*. 2011;72:641–649.

10. Fone DL, Farewell DM, White J, Lyons RA, Dunstan FD. Socioeconomic patterning of excess alcohol consumption and binge drinking: A cross-sectional study of multilevel associations with neighbourhood deprivation. *BMJ Open*. 2013;3:e002337.

11. Halonen JI, Pulakka A, Stenholm S, et al. Change in neighborhood disadvantage and change in smoking behaviors in adults: A longitudinal, within-individual study. *Epidemiology*. 2016;27:803–809.

12. Wilson DK, Kirtland KA, Ainsworth BE, Addy CL. Socioeconomic status and perceptions of access and safety for physical activity. *Ann Behav Med*. 2004;28:20–28.

13. Cano MA, Wetter DW. Socioeconomic status and smoking cessation: Neighborhood context as an underlying mechanism. *Tex Heart Inst J*. 2014;41:309–310.

14. Galster GC. The mechanism (s) of neighbourhood effects: Theory, evidence, and policy implications. In: Van Ham M, Manley D, Bailey N, Maclellan D, eds. *Neighbourhood Effects Research: New Perspectives*. Dordrecht, Netherlands; 2013:23–56.

15. Feldman PJ, Steptoe A. How neighborhoods and physical functioning are related: The roles of neighborhood socioeconomic status, perceived neighborhood strain, and individual health status risk factors. *Ann Behav Med*. 2004;27:91–99.

16. Markowitz FE. Socioeconomic disadvantage and violence: Recent research on culture and neighborhood control as explanatory mechanisms. *Agress Violent Behav*. 2003;8:145–154.

17. Poortinga W. Perceptions of the environment, physical activity, and obesity. *Soc Sci Med*. 2006;63:2835–2846.

18. Shareck M, Ellaway A. Neighbourhood crime and smoking: The role of objective and perceived crime measures. *BMC Public Health*. 2011;11:930.

19. Richardson AS, Troxel WM, Ghosh-Dastidar M, et al. Pathways through which higher neighborhood crime is longitudinally associated with greater body mass index. *Int J Behav Nutr Phys Act*. 2017;14:1–10.

20. Anders MP, Brekenkamp J, Blettner M, Schlehofer B, Bergbeckhoff G. Association between socioeconomic factors and sleep quality in an urban population-based sample in Germany. *Eur J Public Health*. 2014;24:968–973.

21. Ansoleaga E, Castillo-Carniglia A. Associations between social vulnerability, employment conditions and hazardous alcohol consumption in Chile. *Drug Alcohol Rev*. 2013;32:254–261.

22. Romero AJ. Low-income neighborhood barriers and resources for adolescents’ physical activity. *J Adolesc Health*. 2005;36:253–259.

23. Halonen JI, Kivimäki M, Pentti J, et al. Quantifying neighbourhood socioeconomic effects in clustering of behaviour-related risk factors: A multilevel analysis. *PLoS One*. 2012;7:e32937.

24. Moore K, Diez Roux AV, Auchincloss A, et al. Home and work neighbourhood environments in relation to body mass index: The Multi-Ethnic Study of Atherosclerosis (MESA). *J Epidemiol Community Health*. 2013;67:846–853.

25. Virtanen M, Kivimäki M, Pentti J, et al. School neighborhood disadvantage as a predictor of long-term sick leave among teachers: Prospective cohort study. *Am J Epidemiol*. 2010;171:785–792.

26. Virtanen M, Kivimäki M, Elovainio M, Linna A, Pentti J, Vahtera J. Neighbourhood socioeconomic status, health and working conditions of school teachers. *J Epidemiol Community Health*. 2007;61:326–330.

27. Macdonald L. Associations between spatial access to physical activity facilities and frequency of physical activity; how do home and workplace neighbourhoods in West Central Scotland compare? *Int J Health Geogr*. 2019;18:2.

28. Karriker-Jaffe KJ. Areas of disadvantage: A systematic review of effects of area-level socioeconomic status on substance use outcomes. *Drug Alcohol Rev*. 2011;30:84–95.

29. Glenn F, Cody W, Michael M. *Fixed Effects, Random Effects, and Hybrid Models for Causal Analysis*. Dordrecht: Springer Netherlands; 2013.

30. Airaksinen J, Hakulinen C, Pulkki-Råback L, et al. Neighbourhood effects in health behaviours: A test of social causation with repeat-measurement longitudinal data. *Eur J Public Health*. 2016;26:417–421.

31. Magnusson Hanson LL, Leineweber C, Persson V, Hyde M, Theorell T, Westerlund H. Cohort profile: The Swedish Working Environment Survey (arbetsmiljö-2015/2017). Available at https://www.av.se/arbetsmiljоarbetetsmiljоstatistik-officiell-arbetsskadestatistik/arbetsmiljо-2015/2017.

32. SSM Popul Health. 2017;3:48–56.

33. Webb S, Janus M, Duku E, et al. Neighbourhood socioeconomic status indices and early childhood development. *SSM Popul Health*. 2017;3:48–56.

34. Diez Roux AV, Merkin SS, Arnett D, et al. Neighborhood of residence and incidence of coronary heart disease. *N Engl J Med*. 2001;345:99–106.

35. Kim D, Masyn KE, Kawachi I, Laden F, Colditz GA. Neighborhood socioeconomic status and behavioral pathways to risks of colon and rectal cancer in women. *Cancer*. 2010;116:4187–4196.
36. van Vuuren CL, Reijneveld SA, van der Wal MF, Verhoeff AP. Neighborhood socioeconomic deprivation characteristics in child (0–18 years) health studies: A review. Health Place. 2014;29:34–42.

37. Halonen JI, Pulakka A, Vahtera J, et al. Commuting time to work and behaviour-related health: A fixed-effect analysis. Occup Environ Med. 2020;77:77–83.

38. O’Brien CP. The CAGE questionnaire for detection of alcoholism: A remarkably useful but simple tool. JAMA. 2008;300:2054–2056.

39. Nordin M, Åkerstedt T, Nordin S. Psychometric evaluation and normative data for the Karolinska Sleep Questionnaire. Sleep Biol Rhythms. 2013;11:216–226.

40. Halonen JI, Shiri R, Magnusson Hanson LL, Lallukka T. Risk and prognostic factors of low back pain: Repeated population-based cohort study in Sweden. Spine. 2019;44:1248–1255.

41. Prince SA, LeBlanc AG, Colley RC, Saunders TJ. Measurement of sedentary behaviour in population health surveys: A review and recommendations. Peerj. 2017;5:e4130.

42. Statistics Sweden. SOcioekonomisk indelning (SEI). Meddelanden i samordningsfrågor för Sveriges officiella statistik [in Swedish]. 1982:4. Available at https://www.scb.se/hitta-statistik/statistik-efter-amne/ovrigt/ovriga-publikationer-ej-statistik/mis-19824.-socioekonomisk-indelning-sei/.

43. Magnusson Hanson LL, Theorell T, Bech P, et al. Psychosocial working conditions and depressive symptoms among Swedish employees. Int Arch Occup Environ Health. 2009;82:951–960.

44. Magnusson Hanson LL, Westerlund H, Leinenweber C, et al. The Symptom Checklist-core depression (SCL-CD6) scale: Psychometric properties of a brief six item scale for the assessment of depression. Scand J Public Health. 2014;42:82–88.

45. Heikkilä K, Nyberg ST, Fransson EI, et al.; IPD-Work Consortium. Job strain and alcohol intake: A collaborative meta-analysis of individual-participant data from 140,000 men and women. PLoS One. 2012;7:e40101.

46. Karasek Jr RA. Job demands, job decision latitude, and mental strain: Implications for job redesign. Adm Sci Q. 1979;4:285–308.

47. Magnusson Hanson LL, Westerlund H, Chungkham HS, et al. Job strain and loss of healthy life years between ages 50 and 75 by sex and occupational position: Analyses of 64 934 individuals from four prospective cohort studies. Occup Environ Med. 2018;75:486–493.

48. Allison PD. Fixed effects regression methods in SAS. 31st Annual SAS Users Group International Conference. San Francisco, California; 2006.

49. Mallampalli MP, Carter CL. Exploring sex and gender differences in sleep health: A Society for Women’s Health Research Report. J Womens Health (Larchmt). 2014;23:553–562.

50. Wilsnack RW, Wilsnack SC, Kristjanson AF, Vogeltanz-Holm ND, Gmel G. Gender and alcohol consumption: Patterns from the multinationl GENACIS project. Addiction. 2009;104:1487–1500.

51. Statistics Sweden. Concentrations of workplaces outside localities 2010. Available at http://share.scb.se/ov9993/data/publikationer/statistik/ovrigt/ovriga-publikationer-ej-statistik/pong/publikationer/mis-19824.-socioekonomisk-indelning-sei/.

52. Bryden A, Roberts B, Petticrew M, McKee M. A systematic review of the influence of community level social factors on alcohol use. Health Place. 2013;21:70–85.

53. Collins SE. Associations between socioeconomic factors and alcohol outcomes. Alcohol Res. 2016;38:83–94.

54. Cerdá M, Diez-Roux AV, Tchetgen ET, Gordon-Larsen P, Kiefe C. The relationship between neighborhood poverty and alcohol use: Estimation by marginal structural models. Epidemiology. 2010;21:482–489.

55. Brenner AB, Borrell LN, Barrientos-Gutierrez T, Roux AVD. Longitudinal associations of neighborhood socioeconomic characteristics and alcohol availability on drinking: Results from the Multi-Ethnic Study of Atherosclerosis (MESA). Soc Sci Med. 2015;145:17–25.

56. McCusker MT, Basquille J, Khwaja M, Murray-Lyon IM, Catalán J. Hazardous and harmful drinking: A comparison of the AUDIT and CAGE screening questionnaires. QJM. 2002;95:591–595.

57. Martin JK, Roman PM, Blum TC. Job stress, drinking networks, and social support at work: A comprehensive model of employees’ problem drinking behaviors. Sociol Q. 1996;37:579–599.

58. Davis CG, Thake J, Vilhena N. Social desirability biases in self-reported alcohol consumption and harms. Addict Behav. 2010;35:302–311.