FRACTIONAL EULER LIMITS AND THEIR APPLICATIONS*

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Abstract. Generalizations of the classical Euler formula to the setting of fractional calculus are discussed. Compound interest and fractional compound interest serve as motivation. Connections to fractional master equations are highlighted. An application to the Schlögl reactions with Mittag-Leffler waiting times is described.

Key words. Euler limit formula, Mittag-Leffler, master equation

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1. Introduction. Euler’s famous limit formula states, as \( n \to \infty \),

\[
(1 + \frac{1}{n})^n \to \exp(1) = e \approx 2.71828 \ldots .
\]

Euler (Introductio, 1748) was motivated by quandaries such as if a “man borrowed 400,000 florins at the usurious rate of five percent annual interest . . .” [58]. Indeed, the special number \( e \) of calculus and possibly also the formula (1) may have been discovered in this way in 1683, although not by Euler; it was another Swiss mathematician, Bernoulli, in connection to his own studies of compound interest. Others certainly made contributions, including Mercator’s earlier work on the logarithm. Euler’s limit is one way to introduce exponential growth: as the discrete compounding intervals become more frequent they tend to a continuous exponential limit. For clarity, in Euler’s example the principal is \( P = 400.000 \) and the interest rate is \( r = 0.05 \), and if the term of the loan was for one year with \( n \) compounding periods, then the total repaid after one year would be

\[
P \left( 1 + \frac{r}{n} \right)^n.
\]

For example, \( n = 12 \) if compounded monthly, \( n = 365 \) if compounded daily, and the limit \( n \to \infty \) would correspond to compounding continuously.

Such examples of compounding processes are memoryless in the sense that the future is conditionally independent of the past, given the present. Other processes have memory and depend on their history. The fractional calculus [50, 9, 23, 4, 28, 39, 45, 43, 44, 48, 16, 32, 8, 29], unknown to Euler (though his oeuvre is related to its origins), offers a mathematical framework for such processes. That involves a generalization of the derivative to a fractional derivative, alongside which the continuous exponential function is generalized to a continuous Mittag-Leffler function:

\[
E_\alpha(z) = \sum_{k=0}^{\infty} \frac{z^k}{\Gamma(\alpha k + 1)}.
\]

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‡\( E_\alpha(z) \) in (2) is the one-parameter Mittag-Leffler function, whose namesake completed his thesis at Uppsala University, while the two-parameter Mittag-Leffler function, \( E_{\alpha,\beta}(z) \), which we touch on in (33), was introduced by Anders Wiman, who also shares a connection to Uppsala [22, 47, 46].
Table 1
Comparing representations of exponential functions and of Mittag-Leffler functions.

|                      | Exponential $e^t$ | Mittag-Leffler $E_\alpha(t)$ |
|----------------------|-------------------|-------------------------------|
| Taylor series        | $\sum_{n=0}^{\infty} \frac{t^n}{n!}$ | $\sum_{k=0}^{\infty} \frac{t^k}{\Gamma(\alpha k + 1)}$ |
| Cauchy integral      | $\frac{1}{2\pi i} \int_C e^z \frac{1}{z - t} \, dz$ | $\frac{1}{2\pi i} \int_C e^z \frac{z^{\alpha - 1}}{z^\alpha - t} \, dz$ |
| Euler limit          | $\lim_{n \to \infty} \left( 1 + \frac{t}{n} \right)^n$ | ? |

Here $\alpha$ smoothly interpolates between the usual calculus at one extreme ($\alpha = 1$) and an ever "more fractional calculus" toward the other extreme ($\alpha \to 0$). In this article, always $0 \leq \alpha \leq 1$. The familiar power series for the exponential, $E_1(t) = 1 + t/1! + t^2/2! + \cdots$, is recovered when $\alpha = 1$. Fractional processes require specialized numerical algorithms [49, 59] and we use codes of Garrappa [19] for (2). Table 1 collects various representations of these functions. A missing entry suggests a question: What might be the fractional generalization of the discrete limiting process in Euler's famous limit formula? This question will eventually lead us to the resolvent but we begin with more elementary approaches.

1.1. A recursive formulation. Revisit Euler’s formula, recursively:

\[ y_0 = 1, \]
\[ y_j = (1 + h) y_{j-1}. \]

(3)

It is common to allow a time parameter $t$ in Euler’s limit formula (1), which often appears as

\[ \lim_{n \to \infty} \left( 1 + \frac{t}{n} \right)^n = \exp(t). \]

This is accommodated in (3) by setting the step size to be $h = t/n$ for fixed $t$. The limit $n \to \infty$ is the same as $h \to 0$, with $t$ held constant. This recursive formulation certainly has the property that we expect: $y_n \to \exp(t)$ as $n \to \infty$.

In fact, (3) is precisely the Euler method for approximating the simple differential equation $dy/dt = y$ with familiar solution $y(t) = \exp(t)$ when $y(0) = 1$. Finite differences of the continuous equation $dy/dt = y$ lead to the discrete approximation

\[ \frac{y_j - y_{j-1}}{h} = y_{j-1}. \]

(4)

Rearranging yields the Euler formula. Comparing Taylor series shows that the local error over a single time step of size $h$ is $O(h^2)$, but errors accumulate over the many steps it takes to reach $y_n$ so by the last step the global error, $y_n - \exp(t)$, is $O(h)$.

Various interpretations of compound interest are possible. A different candidate, in which the interest rate is not fixed, could come by setting $r_j = t/j$ and replacing (3) with $y_j = y_{j-1}(1 + r_j)$. A solution is possible in terms of special functions, $y_n = y_0 \Gamma(n + 1 + t)\Gamma(n + 1)\Gamma(1 + t)^{-1} = (tB(t, n + 1))^{-1}$ (where $\partial/\partial x B(x, y) = B(x, y)(\Psi(x) - \Psi(x + y))$ and $\Psi(x) = d/dx \ln \Gamma(x) = \Gamma'(x)/\Gamma(x)$ is the digamma function), although such alternatives are not further explored here.
This is first order accuracy: convergence is slow and Euler’s formula is not usually a good numerical choice. Nevertheless, it remains pertinent to compound interest.

Stability is another important property of a numerical method that is related to compound interest. For example, for the stable continuous equation \(dy/dt = -y\) with solution in terms of exponential decay \(y(t) = e^{-t}y(0)\), stability requirements place a restriction on the step size, \(|1 - h| < 1\), of the explicit Euler forward difference construction \(y_j = (1 - h)y_{j-1}\). Stability in relation to backward differences arises later in (10) and (43). Errors accumulate in a way that is analogous to compound interest on a bank account. Errors made early contribute more to the final error. Ensuring numerical stability is tantamount to ensuring that the “compound interest” on those local errors does not grow too usuriously.

1.2. A candidate for a fractional Euler formula. A first candidate answer to our motivating question from the Introduction could be, again recursively,

\[
y_0 = 1,
\]

\[
y_j = (1 + h^\alpha \Gamma(1 - \alpha)) y_{j-1} + \frac{y_{j-2} - y_{j-1}}{2^\alpha} + \cdots + \frac{y_0 - y_1}{j^\alpha}.
\]

Here \(\Gamma(z) = \int_0^\infty x^{z-1} \exp(-x)dx\) is the Gamma function that generalizes the usual factorial function \(n!\) from the integers to complex numbers. To arrive at (5), generalize the simple finite difference construction (4) that led to Euler’s limit. Begin by defining the Caputo fractional derivative \(D^\alpha_t\) of order \(\alpha\) via its action on a smooth function

\[
D^\alpha_t f(t) = \frac{1}{\Gamma(1 - \alpha)} \int_0^t \frac{f'(s)}{(t-s)^\alpha} ds
\]

and introduce a time-fractional generalization of our simple differential equation

\[
D^\alpha_t y = \lambda y
\]

with solution \(y(t) = E_\alpha(\lambda t^\alpha) y(0)\).

This fractional analogue offers a continuous target for a discrete process. Set \(\lambda = 1\) for simplicity and discretize both sides of \(D^\alpha_t y = \lambda y\). By quadrature on the integral in the Caputo derivative, this procedure results in

\[
\frac{1}{\Gamma(1 - \alpha)} h \left( \frac{(y_1 - y_0)/h}{(jh)^\alpha} + \cdots + \frac{(y_j - y_{j-1})/h}{(h)^\alpha} \right) = y_{j-1}.
\]

Rearranging yields the proposal (5). Importantly, (5) converges to the solution of the fractional differential equation in terms of the Mittag-Leffler function. For \(0 < \alpha < 1\), \(y_n \to E_\alpha(t^\alpha)\), as expected.

Figure 1 compares the usual notion of compound interest with what might be named “fractional compound interest.” Concerning this candidate (5), we note the following: (i) The powers of \((1 + h)\) present in the Euler limit are generalized to powers of \((1 + h^\alpha \Gamma(1 - \alpha))\) in (5). (ii) Convergence of the Euler limit is slow, but convergence of the fractional generalization is even slower. Convergence also depends on \(\alpha\).
singularity in the Gamma function $\Gamma(1 - \alpha)$ as $\alpha \to 1$ is one numerical issue—this recursion is not a good numerical method for computing the Mittag-Leffler function.

(iii) The Euler limit (3) is memoryless, requiring only the present value $y_{j-1}$ to advance to the next value of $y_j$, but the fractional generalization (5) requires the whole history of values $y_0, \ldots, y_{j-1}$ in order to advance. (iv) Undergraduate calculus textbooks proceed via logarithms to derive the Euler limit, so a different approach to generalizing the Euler limit, not explored here, could come via a fractional generalization of the logarithm $\log(t)$.

Fractional decay. A fractional generalization of the usual exponential decay process is modeled by $D^\alpha t y = -y(t)$, with solution $y(t) = E_\alpha(-t^\alpha)y(0)$. The same approach to discretization that led to (5) now leads to the recursive formulation: $y_0 = 1$, and

$$y_j = \left(1 - h^\alpha \Gamma(1 - \alpha)\right) y_{j-1} + \frac{y_{j-2} - y_{j-1}}{2^\alpha} + \cdots + \frac{y_0 - y_1}{j^\alpha}.$$  

This is the counterpart to (5) when the argument to the Mittag-Leffler function is negative. Apart from a minus sign in the first term in parentheses, it is otherwise identical with (5): the “memory terms” (connected to a memory function later in (61)) have the same form in both growth and decay versions.

1.3. A Grünwald–Letnikov approach. The binomial term $(1 + h)^n$ that appears in the Euler limit suggests another approach to its generalization could come via the Grünwald–Letnikov (GL) derivative, which defines fractional derivatives via a binomial generalization of a binomial-like expansion. With this in mind, another way to express the same fractional model $D_t^\alpha y = \lambda y$ in (7) is

$$y(t) = y(0) + \lambda I^\alpha y,$$

where $I^\alpha$ is the GL fractional integral operator.\footnote{There are different conventions for defining fractional derivatives, which vary in details of how they handle initial conditions, but there is little distinction in relation to fractional integrals.} (Indeed, one way to formalize the notion of solution to $D^\alpha y = \lambda y$ is as a solution of the Volterra integral equation $y(t) = y(0) + \int_0^t k(t - u)\lambda y(u)du$, where $k(t) = t^{\alpha-1}/\Gamma(\alpha)$ is a memory-like function. This...
integral displays the causal nature of the problem: the solution at the present time can be computed using only information from the past, without needing information from the future.) Hence (9) is another representation of the Mittag-Leffler function.

A Grünwald–Letnikov approach to fractional calculus is attractive because the construction is discrete from the beginning (analogous to the way that the usual calculus begins, as continuous limits of finite difference constructions), unlike other approaches (such as Riemann–Louiville), which begin with continuous integer-order integrals or derivatives. Set \( \lambda = 1, \ h = t/n, \) and replace the continuous integral \( I^\alpha \) in (9) by the discrete GL construction to obtain \( y_n = y_0 + h^\alpha \sum_{j=0}^{n-1} (-1)^j \binom{\alpha}{j} y_{n-j}. \) Solving for \( y_n \) gives another candidate for a fractional Euler formula:

\[
y_n = \frac{1}{(1 - h^\alpha)} \left( y_0 + h^\alpha \sum_{j=0}^{n-1} w_{n-j} y_j \right).
\]

This is an implicit numerical scheme coming from backward differences (revisited later in (43)) in the GL construction, so compared to (5), we expect better stability properties from this GL scheme (10). The weights are \( w_j = (-1)^j \binom{\alpha}{j}. \) Setting \( y_0 = 1, \) (10) does satisfy \( y_n \to E_\alpha(t^\alpha) \), as expected.

2. A Cauchy integral representation. Guiding the search for a discrete construction of a fractional Euler limit has been the principle that it should converge to a continuous Mittag-Leffler function. A discretization of the Cauchy integral representation of the Mittag-Leffler function offers another path to this end.

To get from a series representation of the Mittag-Leffler function (2) to the Cauchy integral representation, start with a Laplace transform of the series for \( E_\alpha(z) \) in (11). There are two cases:

\[
L\{E_\alpha(t^\alpha)\} = \frac{s^{\alpha-1}}{s^\alpha - \lambda} = \frac{1}{s - s^{1-\alpha} \lambda}.
\]

The special case \( \lambda = 1 \) arises often: \( L\{E_\alpha(-t^\alpha)\} = s^{\alpha-1}/(1 + s^\alpha) = (s^{1-\alpha} + s)^{-1}. \) Here the Laplace transform is \( \hat{f}(s) = L\{f(t)\} = \int_0^\infty \exp(-st)f(t)dt, \) and the inverse transform is \( f(t) = L^{-1}\{\hat{f}(s)\} = (2\pi i)^{-1} \int_C \exp(st)\hat{f}(s)ds, \) where the contour \( C \) is a line parallel to the imaginary axis and to the right of all singularities of \( \hat{f} \). The inverse transform gives

\[
E_\alpha(t^\alpha) = L^{-1}\{L\{E_\alpha(t^\alpha)\}\} = \frac{1}{2\pi i} \int_C e^{st} \frac{s^{\alpha-1}}{s^\alpha - \lambda} ds
\]

and after a change of variables this leads to Mittag-Leffler’s representation in Table 1:

\[
E_\alpha(z) = (2\pi i)^{-1} \int_C e^{z s^{\alpha-1}} (s^\alpha - z)^{-1} ds.
\]

The contour \( C \) must enclose all singularities and branch points.

2.1. Revisiting the Cauchy integral on a branch cut. Put \( z = \lambda t^\alpha \) in Table 1 to focus on

\[
E_\alpha(t^\alpha) = \frac{1}{2\pi i} \int_C e^{st} \frac{s^{\alpha-1}}{s^\alpha - \lambda t^\alpha} ds.
\]

By collapsing the contour to the real axis (Figure 2) we will now arrive at another representation, in (13) and (21). There are two cases: \( E_\alpha(+t^\alpha) \) is treated separately from \( E_\alpha(-\lambda t^\alpha) \).
Deforming the contour of Mittag-Leffler’s representation in (12) to collapse the integral to the real axis. Nearby A, there is a pole on the real axis at \( z = t \lambda \). The displayed contour corresponds to the case when the argument to the Mittag-Leffler function is positive and accompanies the representation of \( E_\alpha(+\lambda t^\alpha) \) in (17).

2.1.1 A negative argument: \( E_\alpha(-\lambda t^\alpha) \). This representation is well-known:

\[
E_\alpha(-\lambda t^\alpha) = \int_0^\infty w_-(s) \exp(-st)ds.
\]

Here \( w_-(s) \) is the probability density

\[
w_-(s) \equiv \lambda \frac{\sin(\alpha \pi)}{\pi} \frac{s^{\alpha-1}}{s^{2\alpha} + 2\lambda s^\alpha \cos(\alpha \pi) + \lambda^2} \geq 0.
\]

Equation (13) shows the Mittag-Leffler function as a mixture of exponentials. An example of the weighting, \( w_- \), of the components in that mixture is shown in Figure 3 (left). As \( \alpha \to 1 \) the weighting converges to Dirac’s delta distribution \( \delta(s-\lambda) \) centered at \( \lambda \), so that the “mixture” becomes a pure exponential with rate \( \lambda \), and the Mittag-Leffler function \( E_\alpha(-\lambda t^\alpha) \) smoothly transforms to the exponential function \( \exp(-\lambda t) \).

From (13), \( \mathcal{L} \{ E_\alpha(-\lambda t^\alpha) \} \) is “the Laplace transform of the Laplace transform of \( w_- \):”

\[
\mathcal{L} \{ \mathcal{L} \{ w_-(s) \} \} = \mathcal{L} \{ E_\alpha(-\lambda t^\alpha) \} = \frac{s^{\alpha-1}}{\lambda + s^\alpha}.
\]

2.1.2 A positive argument: \( E_\alpha(+\lambda t^\alpha) \). Although (13) displays two purely real representations, recognizing equality seems to have come historically by passing...
through the imaginary domain, in line with Hadamard’s dictum [35], as we now do again for $E_\alpha(\lambda t^\alpha)$.

The contour integral is the sum of the contributions from the four parts marked $A$, $B$, $C$, and $D$ of the deformation of the contour in Figure 2. There is a pole on the real axis at $z = i\lambda \frac{\pi}{2}$. Deforming the path near $A$ shows this pole contributes a residue of $\exp(t\lambda \frac{\pi}{2})/\alpha$. Combined, the paths marked by $B$ cancel to make zero net contribution. The path nearby the origin marked by $C$ can be deformed to approach a circle that makes zero contribution. The origin is a branch point and the negative real axis is a branch cut. This is because (12) involves a term of the form

$$E \equiv \exp(\log|s| + i\arg(s)),$$

which is analytic on $\mathbb{C}\setminus(-\infty,0]$. The two semi-infinite paths, displayed on either side of the negative real axis, near $D$, do not cancel because of the discontinuity across the negative real axis. For $s$ on the negative real axis, think of $s = x$ and $s^\alpha = (-x)^\alpha \exp(i\pi\alpha)$, where $\arg(s) = +\pi$ on the path coming from above the real axis and $\arg(s) = -\pi$ on the path from below. The two paths contribute:

$$\frac{1}{2\pi i} \int_{-\infty}^{0} (-x)^{\alpha-1} e^{-ix(\alpha-1)} e^{\pi \alpha} e^{\pi} dx + \frac{1}{2\pi i} \int_{0}^{\infty} (-x)^{\alpha-1} e^{ix(\alpha-1)} e^{-\pi \alpha} e^{-\pi} dx.$$

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$$\frac{1}{2\pi i} \int_{-\infty}^{0} (-x)^{\alpha-1} e^{-ix(\alpha-1)} e^{\pi \alpha} e^{\pi} dx + \frac{1}{2\pi i} \int_{0}^{\infty} (-x)^{\alpha-1} e^{ix(\alpha-1)} e^{-\pi \alpha} e^{-\pi} dx.$$

The change of variables $x \rightarrow -x$ converts this to an integral on the positive real axis and manipulation brings that integral to the form $-\int_{0}^{\infty} w_(s) \exp(-st) ds$ with weighting $w_+$ in (16):

$$w_+(s) \equiv -(-\lambda)^{\frac{\sin(\alpha \pi)}{\pi}} \frac{s^{\alpha-1}}{s^{2\alpha} + 2(-\lambda)s^\alpha \cos(\alpha \pi) + (-\lambda)^2} \ge 0.$$

The minus signs in (16) emphasize that $w_+(s)$ can be obtained from $w_-(s)$ in (14) by merely changing the sign of $\lambda$ and by including one more overall sign change. Putting all the pieces together gives

$$E_\alpha(\lambda t^\alpha) = \frac{\exp(t\lambda \frac{\pi}{2})}{\alpha} - \int_{0}^{\infty} w_+(s) \exp(-st) ds.$$

It will be helpful to identify some properties of (17).

Completing the square shows the denominator is positive, and thus for $0 < \alpha < 1$, the weighting function $w_+(s)$ is positive. For moderately large $t$, the dominant contribution to $E_\alpha(\lambda t^\alpha)$ in (17) comes from $\exp(t\lambda \frac{\pi}{2})/\alpha$, while the branch cut integral is the smaller moiety. As $\alpha \rightarrow 1$, the Mittag-Leffler function tends to the exponential so the contribution from the integral in (17) tends to zero.

When numerically integrating either density $w_-(s)$ or $w_+(s)$, a challenge may arise due to the the scaling as $s \rightarrow 0^+$, and due to slow decay as $s \rightarrow \infty$ (Figure 3, left), illustrates this only for $w_-(s)$ but the situation for $w_+(s)$ is very similar. A possible remedy is to let $v(x) = sw_+(s)$ with the substitution $s = \exp(x)$. In the new variables the scalings go exponentially to zero (Figure 3, right): $v(\exp(x)) = O(\exp(-\alpha |x|))$ as $|x| \rightarrow \infty$. For example, we can numerically evaluate the integral on the right of (with $ds = s dx$)

$$0 \le C \equiv \int_{0}^{\infty} w_+(s) ds = \int_{-\infty}^{\infty} v(x) dx,$$
which can reasonably be handled by first truncating to a finite subset of the real line and then using an equally spaced quadrature rule such as the trapezoidal rule.

It transpires that

\[ C = \frac{1}{\alpha} - 1 \]

is a normalization constant for the probability density

\[ W_+(s) \equiv \frac{1}{C}w_+(s) = \frac{\alpha}{1-\alpha}w_+(s). \]

For example, if \( \alpha = 1/2 \), then \( C = 1 \), and comparing (14) and (16), this is a special case where the densities are the same: \( W_+ = w_+ = w_- \). The representation of the Mittag-Leffler function (17) becomes

\[ E_{\alpha}(+\lambda t^{\alpha}) = \frac{\exp(t\lambda^{\frac{1}{\alpha}})}{\alpha} - \frac{1-\alpha}{\alpha} \int_{0}^{\infty} W_+(s) \exp(-st)ds. \]

The last term in (21) is a mixture of exponential distributions

\[ \phi_{W_+}(t) \equiv \int_{0}^{\infty} W_+(s) \exp(-st)ds. \]

Compared to a pure exponential distribution, this \( W_+ \)-mixture has heavier tails, as does the Mittag-Leffler distribution (Figure 4). Thus, collapsing the Cauchy integral to the real axis leads naturally to a density \( W_+ \) associated with a Mittag-Leffler function of a positive argument, analogous to the way that \( w_- \) arises from the Cauchy integral associated with a negative argument. Connections to Wright functions and to Fox H-functions are important but not discussed here [40, 42]. Next, we derive formulas to facilitate sampling from the \( W_+ \) density.

Fig. 4. Comparison of survival times from three distributions. (i) A pure, unit rate exponential distribution. (ii) A Mittag-Leffler distribution, which is the mixture of exponentials in \( \int_{0}^{\infty} w_-(s) \exp(-st)ds \) of (13) with the probability density \( w_-(s) \) of (14). (iii) The mixture of exponentials \( \phi_{W_+}(t) = \int_{0}^{\infty} W_+(s) \exp(-st)ds \) in (22) with density \( W_+ \) as in (20). Parameters: \( \alpha = 0.9 \), \( \lambda = 1 \).
The cumulative distribution function. To study the function $G$ introduced below in (26), we begin by defining the functions

$$g(s) = \int_0^s w_+(\sigma) d\sigma \quad \text{for } 0 < s < \infty$$

and

$$\tilde{g}(s) = \frac{1}{2\alpha \pi} \arctan \left( \frac{-\lambda + s^\alpha - 2\lambda \cos(\alpha \pi)}{\lambda + s^\alpha - 2\lambda \cos(\alpha \pi)} \tan \frac{\alpha \pi}{2} \right) - \frac{1}{2\alpha \pi} \arctan \left( \frac{\lambda + s^\alpha - 2\lambda \cos(\alpha \pi)}{-\lambda + s^\alpha \tan \frac{\alpha \pi}{2}} \right).$$

(23)

The former is continuous but if $0 < \alpha < 1/3$, then the latter is discontinuous at the points

$$s_1 = (2\lambda \cos(\alpha \pi) - \lambda)^{1/\alpha} \quad \text{and} \quad s_2 = \lambda^{1/\alpha};$$

indeed, with $C$ in (19),

$$g(s) = C + \begin{cases} \tilde{g}(s) - \frac{1}{\alpha} & \text{if } 0 < s < s_1, \\ \tilde{g}(s) - \frac{1}{2\alpha} & \text{if } s_1 < s < s_2, \\ \tilde{g}(s) & \text{otherwise } (s_2 < s). \end{cases}$$

(24)

However, if $1/3 < \alpha < 1$, then $s_1$ becomes imaginary and

$$g(s) = C + \begin{cases} \tilde{g}(s) - \frac{1}{\alpha} & \text{if } 0 < s < s_2, \\ \tilde{g}(s) & \text{otherwise } (s_2 < s). \end{cases}$$

(25)

We are allowing $\arctan(\pm \infty) = \pm \pi/2$ at the singularities $s_1$ and $s_2$. Observe that $C \to 0$ as $\alpha \to 1$, and at the other extreme, $C \to \infty$ as $\alpha \to 0$. The larger $C$ is, the “more fractional” the calculus.

At last, with $C = g(\infty)$ as in (18) and (19), define $G : (0, \infty) \to (0, 1)$ by

$$G(T) = \int_0^T W_+(s) ds = \frac{g(T)}{C};$$

(26)

so that $G$ is the cumulative distribution function associated with $W_+$ of (20). As usual, the inverse function $G^{-1} : (0, 1) \to (0, \infty)$ is defined by the property that, when $u = G(T)$, we have $T = G^{-1}(u)$. The inverse is the positive real root

$$T = G^{-1}(u) = (Q(Cu))^{1/\alpha} = Q^{1/\alpha};$$

(27)

where $Q : (0, C) \to (0, \infty)$ is given by

$$Q(v) = \lambda \frac{\sin(\alpha \pi(1 + 2v)) - \sin(\alpha \pi)}{\sin(2\alpha \pi(1 + v))}.$$

(28)

A Laplace transform associated with $W_+$. Let (22) define a survival distribution. The Laplace transform of this survival function is the Laplace transform of the Laplace transform of $W_+$:

$$\phi_{W_+}(t) = \mathcal{L}\{\phi_{W_+}(t)\} = \mathcal{L}\{\mathcal{L}\{W_+(s)\}\}.$$  By (21), $\phi_{W_+}(t) = \alpha(1 - \alpha)^{-1}(\exp(t\lambda^{1/\alpha})/\alpha - E_{\alpha}(+\lambda^{\alpha}))$ so $\mathcal{L}\{\phi_{W_+}(t)\} = \mathcal{L}\{\exp(t\lambda^{1/\alpha})\} - \mathcal{L}\{E_{\alpha}(+\lambda^{\alpha})\} = \frac{\alpha}{1 - \alpha} \frac{1}{\alpha(-\lambda^{1/\alpha} + s)} - \frac{s^{\alpha - 1}}{-\lambda + s^{\alpha}}.$$

(29)
Unlike for the exponential function or for the Mittag-Leffler function, the authors do not know if the corresponding ratio of transforms (significant later in (49)), $\tilde{K}_w(s) = \phi_{W_+}(s)/\phi_{W_+}(s) = \phi_{W_+}(s)/(s\phi_{W_+}(s) - 1)$, has a simple interpretation.

**2.2. A fractional Euler limit, again.** Inspired by the representation of the continuous Mittag-Leffler function as a mixture of exponentials in (13) for a negative argument, or (21) for a positive argument, it is natural to ask the question: Is it possible to write the discrete fractional limit in the form of a weighted sum of regular Euler limits? The answer is “yes” and here are two examples.

Fractional decay, again. Instead of (8), we now propose a different discrete fractional generalization of the Euler formula, namely,

$$
\sum_k w_k(1 - s_k t/n)^n. \tag{30}
$$

Fractional growth, again. Likewise, instead of (5), we also now propose a fractional generalization of the Euler formula in the case of a positive argument, namely,

$$
\left(\frac{1 + (\lambda \pm \beta)/n}{\alpha}\right)^n - \frac{1 - \alpha}{\alpha} \sum_k W_k(1 - s_k t/n)^n. \tag{31}
$$

In both cases of decay and of growth, it is desirable that the proposed sum converges to the corresponding integral for large $n$. That integral is (13) in the case of decay, or (21) in the case of growth. That is, it is desirable to have both the following discrete-to-continuous limits, as $n \to \infty$: (i) $w_k \to w(s)$ and (ii) $\sum_k w_k(1 - s_k t/n)^n \to \int_0^\infty w(s) \exp(-st) ds$. Here the range of $k$ values would depend on $n$, and $w = w_-$ of (14) in the case of decay, or $w = W_+$ of (20) in the case of growth. Always, the discrete weights $w_k = w(s_k)$ integrate to one, i.e., $w_k > 0$ and $1 = \sum_k w_k (s_k - s_{k-1})$. We have merely suggested the general form in (30) and (31)—these formulations can be interpreted as quadrature rules applied to corresponding integrals so there remains the potential for myriad variations in the details not specified here, such as the spacing and the number of grid points $s_k$. These would not usually be good numerical schemes for the reasons highlighted in Figure 3. However, as a weighted sum of regular Euler limits, this form is an especially satisfying fractional generalization.

**2.3. Complete monotonicity.** A smooth function $f$ is completely monotone if all derivatives are monotone: $(-1)^n f^{(n)}(x) \geq 0$. Exponential decay, $e^{-t}$, is the prototypical example. Bernstein’s theorem tells us that all completely monotone functions are representable as a mixture\footnote{A mixture usually refers to a finite sum, or at most countably infinite sum, whereas here in our integral with an exponential kernel we are allowing an uncountably infinite, continuous “mixture” or compound distribution.} of exponentials; being completely monotone is equivalent to being the real Laplace transform of a nonnegative function. For the Mittag-Leffler function with a negative argument, this nonnegative function $w$ is explicitly known in (13) and (14).

The Mittag-Leffler function with a positive argument does not have the completely monotone property. However, the weighting functions $w_+(s)$ in (16) or $W_+$ in (20) are positive so $\phi_{W_+} = \int_0^\infty W_+(s) \exp(-st) ds$, the transform of $W_+$ in (22), does have the property. Thus (21) shows us the Mittag-Leffler function of a positive argument as a combination of two functions, one of which is completely monotone.
Both the exponential and the Mittag-Leffler function are entire functions, and both are completely monotone in the case of a negative argument, and both lose this monotone property in the case of a positive argument. Among our candidates for the fractional generalization of the Euler limit, (30) is more amenable to emphasizing such shared properties of the exponential and the Mittag-Leffler function. As an example, suppose \( t > 0 \) and \( \lambda > 0 \) so for all sufficiently large \( n \), \( 0 < (1 - \lambda t/n) < 1 \). A very useful application of the Euler limit formula \( (1 - \lambda t/n)^n \rightarrow \exp(-\lambda t) \), is to make clear that the exponential of a negative argument satisfies \( 0 < \exp(-\lambda t) < 1 \). The Mittag-Leffler function shares this property:

\[
0 < E_\alpha(-\lambda t^\alpha) < 1.
\]

Representations of the Mittag-Leffler function as a weighted integral of exponentials where the weight is a nonnegative density as in (13), or the discrete fractional Euler limit in (30), make it clear that the Mittag-Leffler function does indeed have this property. In contrast, although the fractional generalizations in (5) or (10) have their own virtues, it is nearly impossible to discern this important property from those formulations. This makes a discretization such as suggested by (30) a more attractive fractional generalization.

3. A probabilistic interpretation.

3.1. Sampling from Mittag-Leffler distributions or the \( W_+ \) mixture. A Mittag-Leffler distribution is characterized by three functions. The survival time is

\[
\phi(t) = E_\alpha(-\lambda t^\alpha)
\]

so the waiting time density \( (\psi(t) = -d\phi(t)/dt = -dE_\alpha(-\lambda t^\alpha)/dt) \) is

\[
(33) \quad \psi(t) = \lambda t^{\alpha-1} E_{\alpha,\alpha}(-\lambda t^\alpha),
\]

where \( E_{\alpha,\beta}(z) \equiv \sum_{k=0}^\infty \frac{z^k}{\Gamma(\alpha+k\beta)} \) is the two-parameter Mittag-Leffler function, and the cumulative distribution function (cdf) \( (1 - \phi(t) = \int_0^t \psi(s)ds) \) is

\[
(34) \quad \text{cdf} (t) = 1 - E_\alpha(-\lambda t^\alpha).
\]

All three functions are nonnegative and, as a probability density, \( 1 = \int_0^\infty \psi(t)dt \).

Putting \( \alpha = 1 \) in these formulas recovers the survival time distribution, cdf, and waiting time density, corresponding to an exponential distribution with parameter \( \lambda \) and mean value \( 1/\lambda \). Unlike the exponential case, (33) shows that the Mittag-Leffler function and hence also the solution of (7) are not differentiable at \( t = 0^+ \). In general care must be taken when differentiating near zero, as happens later in (54).

The survival time (32) is also a mixture of exponentials, \( \phi(t) = E_\alpha(-\lambda t^\alpha) = \int_0^\infty w_-(s) \exp(-st)ds \) ((13), (14)). Such mixtures are amenable to methods for fast simulation of geometric stable distributions involving products of independent random variables \[8, 33, 18, 32, 31\]. Let \( F(T) \equiv \int_0^T w_-(s)ds \) denote the cumulative distribution function of \( w_-(s) \). Given \( u \in (0, 1) \) the inverse function, that solves \( F(T) = u \) for \( T \), is known to be \( T = F^{-1}(u) = \lambda^{\frac{1}{\alpha}} (\sin(\pi \alpha)/\tan(\pi \alpha(1-u)) - \cos(\pi \alpha))^{\frac{1}{\alpha}} \). This permits the known fast sampling procedure for the Mittag-Leffler distribution, via

\[
\tau \leftarrow - \left( \frac{1}{\lambda} \right)^{\frac{1}{\alpha}} \left( \frac{\sin(\pi \alpha)}{\tan(\pi \alpha(1-u_1))} - \cos(\pi \alpha) \right)^{\frac{1}{\alpha}} \log(u_2).
\]
Here \( u_1 \) and \( u_2 \) are independent samples from the uniform distribution on \((0, 1)\). As \( \alpha \to 1 \), (35) reduces to the familiar formula \( \tau \sim -\log(u)/\lambda \) for sampling from an exponential distribution with density \( \lambda \exp(-\lambda t) \).

One way to understand (35) is as the product of two independent random variables, \( Z = XY \), with density \( f_Z(z) = \int_0^\infty f_X(x)f_Y(y) (xy)^{-1} \, dy \). Now read (35) as \( XY \), where \( X \sim (\sin(\pi \alpha)/\tan(\pi \alpha (1 - u_1))) - (\cos(\pi \alpha))^1/2 \) is the inverse transform method for sampling from a density \( f_X(x) = w_{-1}(x) \), and where \( w_{-1} \) is (39) (the special case of the density with \( \lambda = 1 \) in (14), but note that this does not assume \( \lambda = 1 \) in (35)) and where \( Y \sim -\log(u_2)/\lambda^{1/2} \) is the familiar inverse transform method for sampling from an exponential distribution with density \( f_Y = \lambda^{1/2} \exp(-\lambda^{1/2} y) \). Then \( f_Z(z) = \int_0^\infty w_{-1}(x)\lambda^{1/2} \exp(-\lambda^{1/2} x) x^{-1} \, dx \). With the change of variables \( s = 1/x \) this becomes \( f_Z(z) = \int_0^1 w_{-1}(1/s)\lambda^{1/2} \exp(-\lambda^{1/2} s z) s \exp(-s^2) \, ds \). Noticing \( w_-(s) = s^2 w_{-}(1/s) \) and replacing \( z \) by \( t \) this becomes

\[
\int_0^\infty w_{-1}(s) s^{\lambda^{1/2}} \exp(-s^{\lambda^{1/2}} t) \, ds.
\]

This is \(-d/dt \int_0^\infty w_{-1}(s) \exp(-s^{\lambda^{1/2}} t) \, ds = -dE_{\alpha}(-\lambda t^\alpha)/dt = -d\phi(t)/dt \), which is the derivative of the representation of \( \phi(t) \) in (40). Thus (36) is another representation of \( \psi(t) \) because the waiting time density of (33) is always \(-d\phi(t)/dt \). This confirms (35) does indeed sample Mittag-Leffler waiting times.

Instead of the above product form, we could think of the Mittag-Leffler density (36) as a sum of densities of exponential random variables. This suggests a recipe: sample the exponential random variable with parameter \( s^{\lambda^{1/2}} \), where \( s \) is sampled according to the density \( w_{-1}(s) \). The recipe could instead just as well sample the exponential random variable with the parameter in which \( s \) is replaced by \( 1/s \). This is because the same change of variables shows \( \psi(t) = \int_0^\infty w_{-1}(s) s^{\lambda^{1/2}} \exp(-s^{\lambda^{1/2}} t) \, ds = \int_0^\infty w_{-1}(s) s^{1-\lambda^{1/2}} \exp(-\lambda^{1/2} s^{-1}) \, ds \). The sampling formula displayed in (35) corresponds to the latter choice.

This same pattern that works for Mittag-Leffler also works for the survival function \( \phi_{W_+}(t) = \int_0^\infty W_+(s) \exp(-st) \, ds \) (Figure 4, (22)), which is also a mixture of exponentials. Fortunately, we identified both the cumulative distribution of \( W_+ \) in (26) and its inverse in (27) and (28). Thanks to (28) we again have a fast Monte Carlo procedure to sample waiting times from the \( W_+ \) mixture: \( C \leftarrow \frac{1}{\alpha} - 1, v \leftarrow Cu_2, \) and

\[
\tau \leftarrow -\left( \frac{1}{\lambda} \right)^{1/2} \left( \frac{\sin(\alpha \pi (1 + 2v)) - \sin(\alpha \pi)}{\sin(2\alpha \pi (1 + v))} \right)^{1/2} \log(u_2),
\]

where \( u_1, u_2 \) are independent uniform random variables on \((0, 1)\).

### 3.2. Matrix arguments and the special density when \( \lambda = 1 \)

The notation \( w_- \) suppresses the dependence of the density on the parameters \( \alpha \) and \( \lambda \). Put \( \lambda = 1 \) in (13) to obtain

\[
E_\alpha(-t^\alpha) = \int_0^\infty w_{-1}(s) \exp(-st) \, ds,
\]

where \( w_{-1}(s) \) is the special case of the probability density \( w_-(s) \) when \( \lambda = 1 \) in (14):

\[
w_{-1}(s) \equiv \frac{\sin(\alpha \pi)}{\pi} \frac{s^{\alpha - 1}}{s^{2\alpha} + 2s^\alpha \cos(\alpha \pi) + 1} \geq 0.
\]
Make the replacement \( t \leftarrow \lambda^{\frac{1}{\alpha}} t \) (now allowing \( \lambda \) to take any positive value) to obtain

\[
E_\alpha(-\lambda s) = \int_0^\infty w_{-1}(s) \exp(-s \lambda^{\frac{1}{\alpha}} t) ds.
\]

Alternatively, we could arrive at (40) by the change of variables \( s \leftarrow \lambda^{\frac{1}{\alpha}} s \) in (13) and by noticing \( \lambda^{\frac{1}{\alpha}} w_-(\lambda^{\frac{1}{\alpha}} s) = w_{-1}(s) \). Compared to the representation of \( E_\alpha(-\lambda s) \) in (13), here in (40) the weighting density \( w_{-1}(s) \) does not depend on \( \lambda \) (but the exponential \( \exp(st) \) in (13) has been replaced by something that does, \( \exp(-s \lambda^{\frac{1}{\alpha}} t) \), so the dependency has been “moved” from the weighting density to the exponential).

A Mittag-Leffler matrix function. There are various issues when defining a matrix function \[24\]. When defining a Mittag-Leffler function of a matrix \( A \) so the dependency has been “moved” from the weighting density to the exponential, by noticing \( \lambda^{\frac{1}{\alpha}} w_-(\lambda^{\frac{1}{\alpha}} s) = w_{-1}(s) \). Compared to the representation of \( E_\alpha(-\lambda s) \) in (13), here in (40) the weighting density \( w_{-1}(s) \) does not depend on \( \lambda \) (but the exponential \( \exp(st) \) in (13) has been replaced by something that does, \( \exp(-s \lambda^{\frac{1}{\alpha}} t) \), so the dependency has been “moved” from the weighting density to the exponential).

A graph Laplacian matrix. Always in this article a matrix \( A \) has two properties:

i. Off-diagonal entries that are positive or zero: \( a_{ij} \geq 0 \) \( (i \neq j) \).

ii. Diagonal entries that ensure columns sum to zero: \( a_{ii} = -\sum_{i \neq j} a_{ij} \).

When symmetric, this matrix is the familiar graph Laplacian. Many authors use this same terminology for both symmetric and nonsymmetric cases. Such a matrix is an example of a sectorial operator with all eigenvalues in the left half of the complex plane. We often further assume the matrix \( A \) has distinct, negative, and real eigenvalues and one zero eigenvalue. With this notation\(^6\) it is the matrix \((-A)\) that we think of as “positive” (!) so the representation is

\[
E_\alpha(A t^\alpha) = \int_0^\infty w_{-1}(s) \exp(-s(-A)^{\frac{1}{\alpha}} t) ds.
\]

This representation (41) emphasizes the role of \((-A)^{\frac{1}{\alpha}}\), and there are important connections to subordination (allowing a solution for a particular \( \alpha \) to be represented as an integral of the solution corresponding to a different \( \alpha \), for example, and related to the monotone properties of section 2.3) and to powers of operators not discussed here \[27, 60, 1, 6, 21, 12\], [51, Chap. 4].

When \( \alpha = 1 \), \( w_{-1} \) is the Dirac distribution centered at 1 so that (41) is the familiar matrix exponential \( \exp(tA) \). For this class of matrices, \( \exp(tA) \) is a stochastic matrix (columns are probability vectors) associated with a Markov process on discrete states in continuous time. To see that the entries of the matrix exponential are all nonnegative, we could first examine the Taylor series to confirm this at small times, \( t \ll 1 \), and then the semigroup property brings confirmation for all time \( t > 0 \). When \( \alpha \neq 1 \), (41) is the Mittag-Leffler function of the matrix. Notice that \((1, 1, \ldots, 1)A = (0, 0, \ldots, 0)\) and \((1, 1, \ldots, 1)1 = (1, 1, \ldots, 1)\), so multiplication of the Taylor series on the left by a row vector of ones quickly shows that columns of \( E_\alpha(A t^\alpha) \) always sum

\[^6\]A numerical analyst will be frustrated, preferring the opposite sign convention. The reason for this choice of signs is that it is common in the Markov process literature to denote an infinitesimal generator this way, although the usual “Q-matrix” is the transpose of \( A \).
to one. If we could show that all entries of the Mittag-Leffler function of a graph Laplacian matrix are nonnegative (and they are!), then we would know that \( E_\alpha(\mathbb{A} t^\alpha) \) is also a stochastic matrix. Unlike the exponential case, though, we do not have the semigroup property. Nevertheless the Taylor series still confirms nonnegativity for \( \text{small times} \), because \( E_\alpha(\mathbb{A} t^\alpha) = \mathbb{I} + t^\alpha \mathbb{A} + \mathcal{O}(t^{2\alpha}) \), off-diagonals are nonnegative, and for small \( t \) the identity matrix more than compensates for the negative diagonal of \( \mathbb{A} \).

A special \( 2 \times 2 \) matrix example. For \( a, b > 0 \), the diagonalization

\[
\begin{pmatrix}
-a & b \\
 a & -b
\end{pmatrix} =
\begin{pmatrix}
b & 1 \\
0 & a + b
\end{pmatrix}
\begin{pmatrix}
-1 & -1 \\
0 & a + b
\end{pmatrix}
\begin{pmatrix}
a & b \\
-a & b
\end{pmatrix}
\frac{1}{a + b}
\]

shows \( -(\mathbb{A})^{\frac{\alpha}{2}} = (a + b)^{\frac{\alpha}{2} - 1} \) is still a graph Laplacian

\[
-(\mathbb{A})^{\frac{\alpha}{2}} = \begin{pmatrix}
-a & b \\
a & -b
\end{pmatrix}^{\frac{\alpha}{2}} = \begin{pmatrix}
-a & b \\
a & -b
\end{pmatrix} (a + b)^{\frac{\alpha}{2} - 1}.
\]

Analogous to the scalar fractional Euler generalization (30), to accommodate a matrix (30) could become a discretization of (41):

\[
\sum_k w_k \left( \mathbb{I} + \frac{-s_k(\mathbb{A})^{\frac{\alpha}{2}}}{n} t \right)^n.
\]

Here \( w_k = w_{-1}(s_k) \geq 0 \) is a discretization of (39). For all sufficiently large \( n \) the entries of \( (\mathbb{I} - s_k(\mathbb{A})^{\frac{\alpha}{2}} t/n) \) are nonnegative. Thus we see via (42) or (41) that the Mittag-Leffler function is a stochastic matrix. Similar to the numerical method of uniformization, the matrix \( (\mathbb{I} - s_k(\mathbb{A})^{\frac{\alpha}{2}} t/n) \) can be interpreted as a first order approximation of a Markov process in continuous time by a Markov process in discrete time. We can interpret \( E_\alpha(\mathbb{A} t^\alpha) \) (which may itself be a semi-Markov process [52]) as a weighted sum of related Markov processes. Denote (20) with \( \lambda = 1 \) by \( W_{+1}(s) \) and rearrange (21) to

\[
P = \frac{\alpha}{\alpha - \sigma} \left( \frac{\exp((-\mathbb{A})^{\frac{\alpha}{2}} t)}{\sigma} - E_\alpha(-\mathbb{A} t^\alpha) \right) = \int_0^\infty W_{+1}(s) \exp(-s(-\mathbb{A})^{\frac{\alpha}{2}} t) ds.
\]

If \( -(\mathbb{A})^{\frac{\alpha}{2}} \) is also a graph Laplacian, then similarly \( P \) must also be a stochastic matrix.

### 3.3. A Mittag-Leffler function of a graph Laplacian is a stochastic matrix.

Instead of the explicit, forward difference (4) that led to our first candidate, we could replace (4) by an implicit, backward difference of the same continuous equation, \( dy/dt = y \). (In fact the backward differences of the Grunwald–Letnikov approach in section 1.3 are one avenue to the fractional resolvent that we now describe [7].) That leads to the discrete approximation \((y_j - y_{j-1})/h = y_j \). Set \( h = t/n \). Recursively, \( y_n = (1 + t/n)^{-n} \). This is another Euler formula, similar to that in Table 1. Both converge to the same exponential limit,

\[
\lim_{n \to \infty} \left( 1 + \frac{t}{n} \right)^n = \lim_{n \to \infty} \left( 1 - \frac{t}{n} \right)^{-n} = \exp(t).
\]

However, the latter representation is more suggestive of the important connection to the resolvent \( (s\mathbb{I} - \mathbb{A})^{-1} \). This resolvent matrix at \( s \) is defined whenever \( s \) is not an eigenvalue and is involved in solutions of \( dy/dt = \mathbb{A} y \), via an Euler limit

\[
\left( \mathbb{I} - \frac{t}{n} \mathbb{A} \right)^{-n} \longrightarrow \exp(\mathbb{A} t).
\]
Central to the Hille–Yosida and related theorems for semigroups is the relationship between the resolvent and the exponential of the infinitesimal generator $A$. Namely, the resolvent is the Laplace transform of the exponential: $(sI - A)^{-1} = \int_0^\infty \exp(-st) \exp(At)ds$. The Mittag-Leffler functions do not have the semigroup property but the essence of the relationship can be generalized to express the Laplace transform of the Mittag-Leffler function, $E_\alpha(At^\alpha)$, in terms of the resolvent:

\[
(45) \quad s^{\alpha-1}(s^\alpha I - A)^{-1} = \mathcal{L}\{E_\alpha(At^\alpha)\} = \int_0^\infty e^{-st}E_\alpha(At^\alpha)ds.
\]

The scalar version of (45) appears in (11) and very similar steps lead to (45). Thus we can now take the inverse Laplace transform of (45) to express the Mittag-Leffler function in terms of the resolvent. This could lead to what amounts to the same representation as in (12): $E_\alpha(At^\alpha) = 1/2\pi i \int_C e^{s^\alpha}E_\alpha(s^\alpha I - At^\alpha)^{-1}ds$. But a different representation appears by instead using the Post–Widder inversion formula:

\[
(46) \quad \lim_{n \to \infty} \frac{1}{n} \sum_{k=0}^{n-1} b_{k \cdot n+1} (I - (t/n)^\alpha A)^{-k}.
\]

The $b_{k \cdot}$ are the positive constants in the $n$th derivative $d^n/ds^n (s^\alpha I - A)^{-1} = (-1)^n s^{-n-1} \sum_{k=1}^{n+1} b_{k \cdot n+1} (s^\alpha I - A)^{-k}$ that Post–Widder requires.

Representation (46) is a generalization of (44) and in that sense it is yet another fractional generalization of the Euler formula. Although (46) would not usually be a good numerical scheme, it can be usefully applied to affirmatively answer our earlier question concerning nonnegativity of the Mittag-Leffler function of a graph Laplacian, $E_\alpha(At^\alpha)$. First note the pattern of $\pm$ signs of the graph Laplacian $A$ implies the following simple pattern of signs in $(I - (t/n)^\alpha A)$: positive entries on the main diagonal, and negative entries off the main diagonal (although possibly zero entries are allowed). That pattern is displayed here on the left of (47) for the $3 \times 3$ case:

\[
(47) \quad \begin{pmatrix} + & - & - \\ - & + & - \\ - & - & + \end{pmatrix}^{-1} = \begin{pmatrix} + & + & + \\ + & + & + \\ + & + & + \end{pmatrix}.
\]

If a matrix with this pattern of signs has sufficiently large entries on the main diagonal, then all entries of the inverse of that matrix are positive. That “inverse positive” property displayed schematically in (47) is a fact of linear algebra [55], related to the class of $M$-matrices [10], and it has a generalization to operators. Applied to our examples (for all sufficiently large $n$ when the diagonal entries are relatively large enough) the pattern of signs (47) implies that all entries of $(I - (t/n)^\alpha A)^{-1}$ are nonnegative. Powers of a nonnegative matrix are of course nonnegative so $(I - (t/n)^\alpha A)^{-k}$ is likewise nonnegative. Representation (46) now shows that all entries of $E_\alpha(At^\alpha)$ are nonnegative. Having already established unit column sum, we have shown the Mittag-Leffler function of a graph Laplacian, $E_\alpha(At^\alpha)$, is indeed a stochastic matrix.
3.4. Random walks on a graph Laplacian. A continuous time random walk (CTRW) [30] on the set of states \( \{1, 2, \ldots, n\} \) is associated with the \( n \times n \) graph Laplacian matrix \( \Lambda \) as follows. Having just arrived at a state \( j \), the walker waits in that state for an amount of time that is a random variable with a state-dependent probability density \( \psi(j, t) \). We have \( \psi(j, t) > 0 \) and \( 1 = \int_0^\infty \psi(j, t) dt \) because \( \psi \) is a density. The walker then immediately moves to a different state \( i \) with probability

\[
(48) \quad \lambda(i, j) = \frac{a_{ij}}{|a_{jj}|} \quad (i \neq j).
\]

We also define \( \lambda(j, j) = \lambda_{jj} = |a_{jj}| \) (unlike the convention \( \lambda_{jj} = 0 \) of some authors). Notice that the properties of the matrix ensure that \( \lambda(i, j) > 0 \) and \( 1 = \sum_{i \neq j} \lambda(i, j) \).

The memory function. This CTRW is a Markov process if and only if the waiting time in all states is an exponential random variable, so that \( \psi(j, t) = a_{jj} \exp(-a_{jj}t) \).

In that special case the master equation that governs the evolution of the probability \( p(j, t) \) of being in state \( j \) at time \( t \) is \( \frac{\partial}{\partial t} p(t) = \Lambda p(t) \) with solution \( p(t) = \exp(\Lambda t) p(0) \), where \( p(t) \) is a vector with entries \( p(j, t) \). Next, we examine the master equation that governs the probability associated with this CTRW in the more general case that does not make assumptions about the form of the waiting time density—it need not be exponential, for example, in which case the process is not Markov and must therefore exhibit some form of memory. That notion of memory turns out to be made mathematically precise by the ratio of the Laplace transform of the survival function to that of the waiting time density:

\[
(49) \quad \hat{K}(s) = \frac{\hat{\psi}(s)}{\hat{\phi}(s)} = \frac{s\hat{\phi}(s) - 1}{\hat{\phi}(s)}.
\]

Mainardi et al. termed (49) the memory function of the process (although they work with the reciprocal of our notation) [41], and it characterizes the generalized master equation (58) that we describe next.

3.5. A generalized master equation for waiting times that are not exponential. It is difficult to generalize to arbitrary waiting times by working with the probability density description alone. Others have overcome this by a finer description of the process in terms of a flux, and we follow a very similar derivation here. This use of a flux after \( n \) steps is analogous to the way subordinators and random time change representations are used in stochastic process theory to connect absolute time to the random number of steps that occur. It is also analogous to derivations of the classical heat equation that do not simply work with temperature alone, and require a notion of heat flux [16, 53, 3, 14, 34, 51], [30, Chap. 5].

Define \( q_n(j, t) \equiv \text{probability density to arrive at site } j \text{ at time } t, \text{ after exactly } n \text{ steps} \). This nonnegative flux measures the “rate of flow of probability in” to the state. Initially \( (t = 0) \) assume the particle is in state \( j_0 \) and that the particle only just arrived, so \( q_0(j, t) \equiv \delta_{j,j_0}\delta(t) \), where \( \delta \) denotes Dirac’s delta distribution. Sum over all steps to get a total flux \( q(j, t) \equiv \sum_{n=0}^\infty q_n(j, t) \). It is convenient to have notation for the total flux minus the initial flux

\[
(50) \quad q^+(j, t) = \sum_{n=1}^\infty q_n(j, t) = q(j, t) - q_0(j, t) = q(j, t) - \delta_{j,j_0}\delta(t).
\]
Recursively, for \( n \geq 1 \),

\[
q_{n+1}(j, t) = \sum_{i \neq j} \int_0^t q_n(i, u) \Psi(j, t, i, u) du,
\]

where \( \Psi(j, t, i, u) \equiv \) probability density to arrive at site \( j \) at time \( t \) in one step, given it arrived at site \( i \) at time \( u \). Put (51) in (50) and swap order of summation \( q^+(j, t) = \sum_{n=1}^{\infty} (\sum_{i \neq j} \int_0^t q_{n-1}(i, u) \Psi(j, t, i, u) du) = \sum_{n=1}^{\infty} \int_0^t (\sum_{i \neq j} q_{n-1}(i, u)) \Psi(j, t, i, u) du \). If we substitute \( \sum_{n=1}^{\infty} q_{n-1}(i, u) = \sum_{n=0}^{\infty} q_n(i, u) = q(i, u) \) and assume the functional form \( \Psi(j, t, i, u) = \lambda(j, i) \psi(i, t - u) \), where \( \psi(i, t) \equiv \) probability density to leave site \( i \) at time \( t \), given it arrived at site \( i \) at \( t = 0 \) is the waiting time, then

\[
q^+(j, t) = \sum_{i \neq j} \lambda(j, i) \int_0^t q(i, u) \psi(i, t - u) du.
\]

The probability density \( p(j, t) \) to be at state \( j \) at time \( t \), is related to the flux by

\[
p(j, t) = \int_0^t \phi(j, t - u) q(j, u) du,
\]

where the survival time \( \phi(j, t) \equiv \) the probability to remain at site \( j \) for all times in \((0, t)\), having just arrived at \( t = 0 \). Using (50) gives \( p(j, t) = \delta_{j,j_0} \phi(j, t) + \int_{0^+}^t \phi(j, t - u) q^+(j, u) du \), where the integral with lower limit \( 0^+ \) means \( \lim_{\epsilon \to 0^+} \int_\epsilon^t \).

Differentiating,

\[
\frac{d}{dt} p(j, t) = \delta_{j,j_0} \frac{d}{dt} \phi(j, t) + \frac{d}{dt} \int_{0^+}^t \phi(j, t - u) q^+(j, u) du
\]

\[
= \delta_{j,j_0} (-\psi(j, t)) + \phi(j, t - t) q^+(j, t) + \int_{0^+}^t (-\psi(j, t - u)) q^+(j, u) du
\]

\[
= - \int_{0^+}^t \psi(j, t - u) q(j, u) du + \sum_{i \neq j} \lambda(j, i) \int_0^t q(i, u) \psi(i, t - u) du.
\]

Note \( \psi(j, t - t) = \phi(j, 0) = 1 \). We also used (50), (52), and the multivariate chain rule to differentiate the integral.

We now have two equations involving the probability density: (53) and (54). Take the Laplace transform of both. The first equation, (53), is a convolution so the transform is a product of transforms: \( \hat{p}(j, s) = \hat{\phi}(j, s) \hat{\psi}(j, s) \). Multiply by \( \hat{\psi}(j, s) \) and divide by \( \hat{\phi}(j, s) \) to get

\[
\hat{q}(j, s) \hat{\psi}(j, s) = \frac{\hat{\psi}(j, s)}{\hat{\phi}(j, s)} \hat{p}(j, s) = \hat{K}(j, s) \hat{p}(j, s),
\]

where the state-dependent version of the memory function (49) now appears,

\[
\hat{K}(j, s) \equiv \frac{\hat{\psi}(j, s)}{\hat{\phi}(j, s)}.
\]

The second equation (54) involves a derivative on the left and convolutions on the right so the Laplace transform is

\[
s\hat{p}(j, s) - p(j, 0) = -\hat{\psi}(j, s) \hat{q}(j, s) + \sum_{i \neq j} \lambda(j, i) \hat{\psi}(i, s) \hat{q}(i, s).
\]
Use (55) and (56) to replace \( \hat{\psi}(i, s) \hat{q}(i, s) \) by \( \hat{K}(i, s) \hat{p}(i, s) \):

\[
(57) \quad s\hat{p}(j, s) - p(j, 0) = -\hat{K}(j, s) \hat{p}(j, s) + \sum_{i \neq j} \lambda(j, i) \hat{K}(i, s) \hat{p}(i, s).
\]

Take inverse Laplace transforms to finally arrive at the desired master equation

\[
(58) \quad \frac{d}{dt} p(j, t) = -\int_{0}^{t} K(j, t-u) p(j, u) du + \sum_{i \neq j} \lambda(j, i) \int_{0}^{t} K(i, t-u) p(i, u) du.
\]

This master equation does not assume exponential waiting times. The waiting times may have different functional forms in different states. In the special case of exponential waiting times, \( \hat{K}(j, s) = \lambda_{jj} \) in (57) so the \( K(j, t) \) appearing in the convolutions in (58) are Dirac delta distributions and (58) collapses to the usual master equation \( \frac{d}{dt} p(t) = A p(t) \).

### 3.6. Mittag-Leffler waiting times have Mittag-Leffler matrix functions for master equations.

Now specialize the general master equation (58) to Mittag-Leffler waiting times: state \( j \) has survival function \( \hat{\phi}(j, t) = E_{\alpha}(-\lambda_{jj} t^\alpha) \) [25]. The following steps are very similar to [41], although here we work with a matrix. By (15),

\[
\hat{\phi}(j, s) = \mathcal{L}\{w_-(s)\} = \mathcal{L}\{E_{\alpha}(-\lambda_{jj} t^\alpha)\} = \frac{s^{\alpha-1}}{\lambda_{jj} + s^\alpha}.
\]

The transform of its derivative \( \hat{\psi} = -\frac{d}{dt} \hat{\phi} \) is \( s \hat{\phi} - 1 \) so \( \hat{\psi}(j, s) = \frac{\lambda_{jj}}{\lambda_{jj} + s^\alpha} \). The memory function (56) is thus

\[
\hat{K}(j, s) = \lambda_{jj} s^{1-\alpha}.
\]

Substitute into (57):

\[
(59) \quad s\hat{p}(j, s) - p(j, 0) = -\lambda_{jj} s^{1-\alpha} \hat{p}(j, s) + \sum_{i \neq j} \lambda(j, i) \lambda_{ii} s^{1-\alpha} \hat{p}(i, s).
\]

Then divide both sides by \( s^{1-\alpha} \), so

\[
(59) \quad s^{\alpha} \hat{p}(j, s) - s^{\alpha-1} p(j, 0) = -\lambda_{jj} \hat{p}(j, s) + \sum_{i \neq j} \lambda(j, i) \lambda_{ii} \hat{p}(i, s).
\]

The right side is precisely the matrix-vector product \( \Lambda \hat{p} \) (recalling (48)). We recognize the left as the Laplace transform of a Caputo fractional derivative:

\[
\mathcal{L}\left\{ \frac{d^\alpha}{dt^\alpha} f(t) \right\} = s^\alpha f(s) - s^{\alpha-1} f(0^+).
\]

Thus, after an inverse transform, the master equation is \( \frac{d^\alpha}{dt^\alpha} p(t) = \Lambda p(t) \). We may also write this as the matrix-vector version of the scalar model fractional equation (7) from the Introduction (although (60) corresponds to the fractional decay case)

\[
(60) \quad D_t^\alpha p = \Lambda p \quad \text{with solution} \quad p(t) = E_{\alpha}(\Lambda t^\alpha)p(0).
\]

### Conclusion:

Mittag-Leffler waiting times go together with a Mittag-Leffler function of the (graph Laplacian) matrix. This is also a second proof of our earlier observation near (47) that a Mittag-Leffler function of a graph Laplacian is a stochastic matrix.
3.7. A remarkable class of nonnegative matrix functions. While (58) boasts generality, it does not make apparent the simple structure of important special classes of solutions. Let us now elucidate one such special class, in terms of matrix functions. The algebra that permits the simplified form of (59) and thus (60) is the old idea of separation of variables: \( \hat{K}(j, s) = \hat{k}(s) g(j) \). The generalized master equation then reads

\[
\frac{1}{\hat{k}(s)} (s \hat{p}(j, s) - p(j, 0)) = -g(j) \hat{p}(j, s) + \sum_{i \neq j} \lambda(j, i) g(i) \hat{p}(i, s).
\]

The right is \( A \hat{p} \) where the main diagonal entries are \( a_{jj} = -g(j) \). In the time domain, the left is

\[
(\int_0^t \frac{1}{k(t - u)} \, du)
\]

The Caputo derivative (6) corresponds to a well-understood example

\[
k(t) \equiv t^\alpha \Gamma(1 - \alpha),
\]

and the memory terms appearing in our first candidate for the fractional Euler formula (5) can now be understood as coming from a discretization of the corresponding convolution (61) with that particular form of \( k \).

Solutions are matrix functions, computed as the inverse Laplace transform of

\[
\frac{1}{\hat{k}(s)} \left( \frac{s \mathbb{I}}{\hat{k}(s)} - A \right)^{-1}.
\]

Numerical inversion of (63) is well suited to Cauchy contour integral methods, although further research is required to address issues associated with pseudospectra of graph Laplacians [37, 57] and to tailor those methods to various forms of the memory function \( k \). As probabilities, these solutions are nonnegative, so exploring separable forms of the memory function gives rise to a large class of nonnegative matrix functions. Indeed, these functions output stochastic matrices. Confirming unit column sum might again come by multiplying a Taylor series in powers of \( A \) on the left by a row vector of ones. That the first term in such a series should be simply \( \mathbb{I} \) might be established by applying the usual Laplace transform rule that \( f(0^+) = \lim_{s \to \infty} s \hat{f}(s) \) to (63). More interestingly, the very special nonnegativity property is explained as variations of the algebraic structure elucidated in (46) and (47). Inversion of the Laplace transform (63) could express the solution in terms of a resolvent, and then the key observation displayed in (47) is that our graph Laplacian is an example of a resolvent positive operator [2, 5]: for large positive \( s \), all entries of the resolvent matrix \( (s \mathbb{I} - A)^{-1} \) are nonnegative. The resolvent shares the completely monotone properties of section 2.3, and the usual Euler formula in terms of the resolvent (44) quickly shows solutions of the corresponding Cauchy problem preserve positivity. Further exploration of fractional Euler formulas (as variations of (42) or (46)), and further exploration of algebraic structures of master equations including Toeplitz or Hankel structures and other splittings [56, 54], therefore seems promising.
4. Application to the Schlögl reactions. The Schögl model consists of four reactions:

1. \( B_1 + 2X \rightarrow 3X, k_1 = 3 \times 10^{-7}, \)
2. \( B_1 + 2X \leftrightarrow 3X, k_2 = 1 \times 10^{-4}, \)
3. \( B_2 \rightarrow X, k_3 = 1 \times 10^{-3}, \)
4. \( B_2 \leftrightarrow X, k_4 = 3.5. \)

Here \( B_1 = 1 \times 10^5 \) and \( B_2 = 2 \times 10^5 \) are constants that model buffered species. We choose the initial condition \( X(0) = 247 \), which in the deterministic version of the model lies on the separatrix between the two basins of attraction of the lower steady state (at about 85) and of the higher steady state (at about 565). The stochastic model exhibits a bimodal distribution (Figure 5). Fedotov also uses this model as an example, in a physically comprehensive discussion [16].

Such reactions can be modeled in the framework of the chemical master equation, which is a memoryless Markov process with exponential waiting times [20, 15]. Waiting times different from the usual exponential choice could model some form of anomalous diffusion, although there are many issues associated with a careful physical interpretation [16, 11]. Operationally, experimentation with different waiting times comes by simply changing the simulation of the waiting time in the usual Gillespie stochastic simulation algorithm, while keeping all other steps in the algorithm the same [20]. As an example, Figure 5 incorporates Mittag-Leffler waiting times \((\alpha = 0.7)\) in simulations of the Schögl reactions, so the generalized master equation is (60). The waiting time between reactions is simulated via (35):

\[
\tau \leftarrow - \left( \frac{1}{a} \right)^{\frac{1}{\alpha}} \left( \frac{\sin(\pi \alpha)}{\tan(\pi \alpha(1 - u_1))} - \cos(\pi \alpha) \right)^{\frac{1}{\alpha}} \log(u_2),
\]

where \( u_1, u_2 \) are drawn from independent uniform random variables on \((0, 1)\), and where \( a \) is the usual sum of the propensities of the reactions in the current state.

For this model there are some numerical issues associated with truncation to a finite state space [38] and with stationary distributions [13], although numerical experiments indicate good accuracy here. The Mittag-Leffler waiting times manifest themselves in the very long pauses between consecutive reactions (Figure 5, right). With
the usual exponential waiting times the solution of the associated master equation
is only bimodal, but the Mittag-Leffler solution exhibits a third mode at \( \approx 247 \) that
persists for a very long time (Figure 5).

The generalized master equation (58) invites further experimentation with different
waiting times leading to processes with memory. One form of memory could come
from the \( W_+ \) mixture in (22). A fast Monte Carlo method to sample from that \( W_+ \)
mixture is provided in (37). The generalized master equation would be (58) with a
particular memory function \( \hat{K}_W(s) = \hat{\phi}_W(s)/(s\hat{\phi}_W(s) - 1) \), with transform \( \hat{\phi}_W \)
available in (29).

Discussion. Our fractional usury was motivated in part by a quest for the “right”
notion of fractional compound interest, a subject that deserves more attention. With
this in mind, various candidates for a fractional generalization of the closely related
Euler formula have been discussed. This led to a list of three candidates to replace the
missing entry (denoted by a question mark) in Table 1: (5), (10), and (30)–(31). These
were based, respectively, on generalizing the discrete construction that leads to the
usual Euler limit (5), a Grünwald–Letnikov approach (10), or a Cauchy integral (30)–
(31). That all candidates be discrete and converge to the continuous Mittag-Leffler
function has served as a guiding principle. Both a Cauchy-like formulation (30)–(31)
and a formulation in terms of the resolvent (46) are attractive because they have the
satisfying form of being a weighted sum of regular Euler formulas. Together with the
observation that the graph Laplacian is a resolvent positive operator, with pattern of
signs as in (47), one application of these fractional Euler formulas is to show that the
Mittag-Leffler function of a graph Laplacian is a stochastic matrix. Generalizations
of other exponential properties to the Mittag-Leffler setting are destined. For instance,
the Golden–Thompson inequality [17] states that, for Hermitian matrices \( A \) and \( B \),
\[ \text{trace} \left( \exp(A + B) \right) \leq \text{trace} \left( \exp(A) \exp(B) \right) \]. The inequality fails when the exponential
function is replaced by the Mittag-Leffler function, as a scalar example quickly
shows \( E_{1/2}(1 + 1) = \exp((1 + 1)^2) = 109 \approx 5 \times 5 \approx E_{1/2}(1)^2 \), but less naïve generalizations might be possible. Finally, continuing this theme, we
have elucidated connections to master equations that generalize the usual exponential
waiting times to Mittag-Leffler waiting times, exemplified by applications to modeling
and simulation of chemical reactions.

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