A cumulant approach for the first-passage-time problem of
the Feller square-root process

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The paper focuses on an approximation of the first passage time proba-
bility density function of a Feller stochastic process by using cumulants
and a Laguerre-Gamma polynomial approximation. The feasibility of
the method relies on closed form formulae for cumulants and moments
recovered from the Laplace transform of the probability density func-
tion and using the algebra of formal power series. To improve the
approximation, sufficient conditions on cumulants are stated. The re-
sulting procedure is made easier to implement by the symbolic calculus
and a rational choice of the polynomial degree depending on skewness,
kurtosis and hyperskewness. Some case-studies coming from neuronal
and financial fields show the goodness of the approximation even for a
low number of terms. Open problems are addressed at the end of the
paper.

Keywords: hitting times, CIR model, Laguerre series, formal power
series, symbolic calculus.

1. INTRODUCTION

One-dimensional diffusion processes play a key role in the description of
fluctuating phenomena belonging to different fields of applications as physics,
biology, neuroscience, finance and others (Karlin and Taylor, 1981; Øksendal,
1998). In particular, the class of stochastic processes with a linear drift and
driven by a Wiener process is widely used for its mathematical tractability and
flexibility. These models are described by a stochastic differential equation of
the following type

\[ dY_t = (-\tau Y_t + \mu) \, dt + \Sigma(Y_t) \, dW_t, \quad Y_0 = y_0, \]  

(1)

where \( \{W_t\}_{t \geq 0} \) is a standard Wiener process, \( Y_0 = y_0 \) is the initial condition,
\( \tau > 0, \mu \in \mathbb{R} \) and the volatility \( \Sigma(Y_t) > 0 \) are such that a strong solution of
Eq.(1) exists (Arnold (2013) p.105).

The volatility \( \Sigma(Y_t) \) determines the amplitude of the noise and, according
to its dependence on \( Y_t \), it characterizes families of stochastic processes which
are solution of Eq. (1). If

\[ \Sigma(Y_t) = \sqrt{aY_t^2 + bY_t + c}, \quad a, b, c \in \mathbb{R}, \]  

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the solution of Eq. (1) is called Pearson diffusion process (Forman and Sørensen, 2008). The coefficients $a, b$ and $c$ are such that the square root is defined for all the values of the state space $(B_1, B_2)$ of $Y_t$, with $-\infty \leq B_1 < y_0 < B_2 \leq +\infty$. A wide range of well-known processes belongs to this class ($\sigma > 0$):

Ornstein-Uhlenbeck process: $a = b = 0, c = \sigma^2$ and $\Sigma(Y_t) = \sigma$;

Inhomogeneous geometric Brownian motion: $a = \sigma^2, b = c = 0$ and $\Sigma(Y_t) = \sigma Y_t$;

Jacobi diffusion: $a = -\sigma^2, b = \sigma^2$ and $\Sigma(Y_t) = \sigma \sqrt{Y_t(1 - Y_t)} + c$;

Feller process (CIR model): $b = \sigma^2, a = 0$ and $\Sigma(Y_t) = \sigma \sqrt{Y_t} + c$.

Throughout this paper we will focus on this last process for its variety of applications not only in a biological context (Ditlevsen and Lansky, 2006; Feller, 1951; Lansky, Sacerdote, and Tomassetti, 1995) but also in survival analysis, in the modeling of nitrous oxide emission from soil and in other applications such as physics and computer science (see Ditlevsen and Lansky (2006) and references therein). In the mathematical finance it is known under the name of Cox-Ingersoll-Ross model (CIR) (Cox, Ingersoll, and Ross, 1985).

While general properties of the Feller process are well known since long, less known are properties related to first-passage-time (FPT) events which are very significant phenomena in all of the above mentioned situations. In this paper we consider the dynamics of $Y_t$ until it crosses a threshold $S$ for the first time, the so called (upcrossing) FPT, defined as

$$T := \inf\{t \geq 0 : Y_t \geq S|0 < y_0 < S\}. \quad (3)$$

Many contributions in the literature (Giorno et al., 1986; Going-Jaeschke and Yor, 2003; Linetsky, 2004; Masoliver and Perelló, 2014) focus on computing the Laplace transform (LT) of the probability density function (PDF) $g(t) := g(t|y_0, S)$ of $T$, namely

$$\tilde{g}(z) = \int_0^\infty e^{-zt} g(t) \, dt, \quad z > 0. \quad (4)$$

The reason why the literature is focused on the LT of $g(t)$, is that the PDF is usually not known analytically and neither can be obtained by direct inversion of Eq. (4). Nevertheless from $\tilde{g}$ we can compute the probability of crossing the threshold $S$, $\mathbb{P}(T|y_0) = \int_0^\infty g(t) \, dt$, and the mean FPT, $\mathbb{E}[T]$ as follows:

$$\mathbb{P}(T|y_0) = \tilde{g}(z) \bigg|_{z=0} \quad \text{and} \quad \mathbb{E}[T] = -\frac{d\tilde{g}(z)}{dz} \bigg|_{z=0}. \quad (5)$$

Moments of $T$ of any orders can be computed using higher derivatives of $\tilde{g}$, when they exist. As it is well-known, the moments of $T$ allow nice interpretation of statistical properties of the PDF $g(t)$ and of FPT events. A different
strategy to get the moments of $T$ is using the transition PDF of the process $f(y, t | y_0, \tau) = \frac{\partial}{\partial y} \mathbb{P}(Y_t < y | Y_\tau = y_0)$. Indeed if $Y_t$ admits a stationary distribution $W(y) := \lim_{t \to \infty} f(y, t | y_0, 0)$ independent of $y_0$, the Siegert formula (Siegert, 1951) allows us to compute the moments of $T$ as

$$E[T^n] = n \int_{x_0}^{x_S} \frac{2dz}{\sum(Y_t)^2} W(z) \int_{-\infty}^{x} W(x) E[T^{n-1}] dx, \quad n = 1, 2, \ldots$$

Both the depicted strategies are impractical to compute the moments of $T$ for a Feller process. Despite the closed form formula of $\tilde{g}(z)$ (see Section 2), the computation of higher derivatives is awkward and some efforts have focused in evaluating just the mean and the variance of $T$ (Ditlevsen and Lansky, 2006; D’Onofrio, Lansky, and Pirozzi, 2018) or at most the third moment (Giorno et al., 1988). In terms of computational complexity, similar difficulties apply in computing moments of $T$ through Eq. (6), although the stationary distribution of $Y_t$ is known to be a shifted gamma distribution (see Section 2). As the distribution of $T$ is often unavailable, simulations of the paths through Monte Carlo methods are still an efficient tool to get manageable estimations of $g(t)$, useful to analyze especially asymptotic properties. One more strategy consists in writing the FPT distribution as a Sturm-Liouville eigenfunction expansion series, first given for the Feller process in Linetsky (2004), using the classical argument of Kent (1980) and Kent (1982). Although this strategy provides an expression for the FPT density, information on the moments of $T$ can be obtained only numerically and refers exclusively to diffusion processes without natural boundaries. A discussion on FPT of the Feller process in the presence of entrance, exit and reflecting boundary at the origin is given in Martin, Behn, and Germano (2011), solving the Sturm-Liouville boundary problem in the case $\tau = 0$.

The goal of this paper is twofold: to give closed form formulae for the cumulants of $T$ of any order for the Feller process regardless of the nature of the boundaries and to give approximations of $g(t)$ by using moments recovered from cumulants.

Recall that if $T$ has moment generating function $\mathbb{E}[e^{zT}] < \infty$ for all $z$ in an open interval about 0, then its cumulants $\{c_k(T)\}_{k \geq 0}$ are such that

$$\sum_{k \geq 1} c_k(T) \frac{z^k}{k!} = \log \mathbb{E}[e^{zT}]$$

for all $z$ in some (possibly smaller) open interval about 0. Cumulants have nice properties compared with moments such as the semi-invariance and the additivity (McCullagh, 1987). Further properties on cumulants are given in Section 3. Overdispersion and underdispersion as well as asymmetry and tailedness of
the FPT PDF might be analyzed through the first four cumulants. Examples on how to employ the first four cumulants in the estimation of the parameters of a model fitted to data is given in (Antunes et al., 2020; Seneta, 2004).

The employment of cumulants in the FPT literature is not new (Ramos-Alarcón and Kontorovich, 2013). However, their application has been limited to few cases and not in the direction addressed in this paper. Here, the idea to use cumulants essentially relies on the form of $\tilde{g}(z)$ for the Feller process. Indeed $\tilde{g}(z)$ is the ratio of two power series whose algebra is simplified if we consider $\log\tilde{g}(z)$. We take advantage of the formal power series algebra (Charalambides, 2002) to give first a closed form expression of $\{c_k(T)\}$ and then to recover moments. In Section 4, we propose to use a Laguerre series to approximate the PDF $g(t)$ taking into account the properties of $\tilde{g}(z)$. To the best of our knowledge, this approach in evaluating the FPT PDF of the Feller process has not been investigated before in the literature. Such an approximation works if moments (or cumulants) of $T$ are known and gives better results when the series is of Laguerre-Fourier type. As the PDF $g(t)$ is unknown, we give sufficient conditions on the cumulants of $T$ to guarantee the approximation with the Laguerre-Fourier series. We show how to take advantage of the formal power series algebra and of the symbolic calculus (Di Nardo, 2012) in implementing the proposed procedure. Some new results on the Kummer’s function are also given.

Then we apply our method to different case-studies inspired by neuronal and financial models. One of the advantages of the method is that few terms are sufficient to have a good description of $g(t)$ and the complexity of the overall computation is strongly reduced. Statistical arguments motivate the choice of stopping the Laguerre series at the fifth term. The case-studies show that the resulting approximation is accurate also when the sufficient conditions are not completely fulfilled. A discussion section ends the paper, addressing future research and open problems.

2. THE FELLER PROCESS AND THE FPT PROBLEM

We consider model (1) such that the function $\Sigma(\cdot)$ depends on the process itself and on $c \leq 0$. The Feller process investigated here is given by

$$dY_t = (-\tau Y_t + \mu) \, dt + \sigma \sqrt{Y_t - c} \, dW(t).$$

The state space of the process is the interval $(c, +\infty)$. The endpoints $c$ and $\infty$ can or cannot be reached in a finite time depending on the underlying parameters. According to the Feller classification of boundaries (Karlin and Taylor, 1981), $c$ is an entrance boundary if it cannot be reached by $Y_t$ in finite time, and there is no probability flow to the outside of the interval $(c, +\infty)$, that is, the process stays in $[c, +\infty)$ with probability 1. In particular, set $s := 2(\mu - c\tau)/\sigma^2$. Then $c$ is an entrance boundary if $s \geq 1$.

In the absence of a threshold, the Feller process admits a stationary distribution which is a shifted gamma distribution with the following shape, scale
and location parameters

\[ Y_\infty \sim \text{Gamma} \left( s, \frac{1}{2}, \frac{\sigma^2}{\tau}, c \right). \]  

Let \( Y_t \) evolve in the presence of a threshold \( S \). Let \( T \) be the FPT random variable of \( Y_t \) through \( S \) defined in Eq. (3). Three distinct situations for the FPT can occur. Indeed, the process is said to be in the suprathreshold, subthreshold and threshold regimes if \( \mathbb{E}[Y_\infty] > S, \mathbb{E}[Y_\infty] < S \) and \( \mathbb{E}[Y_\infty] = S \), respectively, where the asymptotic mean of \( Y_t \) is

\[ \mathbb{E}[Y_\infty] = \lim_{t \to +\infty} \mathbb{E}[Y_t|y_0] = \frac{\mu}{\tau}. \]  

The Siegert equation (Masoliver and Perell\`a, 2014)

\[ \frac{1}{2} \sigma^2 (y_0 - c) \frac{\partial^2 \tilde{g}(z)}{\partial y_0^2} - (\tau y_0 + \mu) \frac{\partial \tilde{g}(z)}{\partial y_0} - z \tilde{g}(z) = 0 \]  

with initial conditions \( \tilde{g}(z) = 1 \) if \( y_0 \equiv S \) and \( \tilde{g}(z) < +\infty \) for any \( y_0 \), provides the LT of the FPT PDF. Indeed the solution of Eq. (11) is

\[ \tilde{g}(z) = \Phi\left( \frac{z}{\tau}, s, \frac{2\tau(y_0-c)}{\sigma^2} \right) \Phi\left( \frac{z}{\tau}, s, \frac{2\tau(S-c)}{\sigma^2} \right), \quad z > 0 \]  

where \( \Phi \) is the confluent hypergeometric function of the first kind (or Kummer’s function) \( \Phi(a,b;z) = \text{_{1}F_{1}}(a;b;z) \), and

\[ \text{_{p}F_{q}}(a_1,\ldots,a_p;b_1,\ldots,b_q;z) := \sum_{n \geq 0} \frac{\langle a_1 \rangle_n \cdots \langle a_p \rangle_n z^n}{\langle b_1 \rangle_n \cdots \langle b_q \rangle_n n!} \]  

is the generalized hypergeometric function, with \( \langle a \rangle_n = a(a+1)\cdots(a+n-1), n \in \mathbb{N} \) the rising factorial and \( \langle a \rangle_0 = 1 \). For more details on Eqs. (9)-(12) see D’Onofrio, Lansky, and Pirozzi (2018). In particular, the mean of \( T \) is (Giorno et al., 1988)

\[ \mathbb{E}[T] = \frac{(S - y_0)}{\mu - \tau c} + \frac{1}{\tau} \sum_{n \geq 2} \frac{s^n \Gamma(s) \left[ (S - c)^n - (y_0 - c)^n \right]}{n \Gamma(s + n) \left( \frac{\mu}{\tau} - c \right)^n}, \]  

where \( \Gamma(z) = \int_0^\infty x^{z-1} e^{-x} \, dx \) is the gamma function.

3. FPT CUMULANTS

Suppose \( \tilde{g} \) a formal power series (Charalambides, 2002)

\[ \tilde{g}(z) = \sum_{k \geq 0} \tilde{g}_k \frac{z^k}{k!} \in \mathbb{R}[[z]] \]  

is the LT of the FPT PDF.
where \( \mathbb{R}[[z]] \) denotes the ring of formal power series with coefficients in \( \mathbb{R} \). Then \( \log \tilde{g}(z) \) is well defined

\[
\log \tilde{g}(z) = \sum_{k \geq 1} c_k \frac{z^k}{k!} \tag{16}
\]

and the coefficients \( \{c_k\}_{k \geq 1} \) are named formal cumulants of \( \{\tilde{g}_k\}_{k \geq 0} \). There are different formulae expressing formal cumulants in terms of \( \{\tilde{g}_k\}_{k \geq 0} \), Di Nardo (2012). Here we use the logarithmic (partition) polynomials \( \{P_k\} \) such that

\[
c_k = P_k(\tilde{g}_1, \ldots, \tilde{g}_k), \quad k \geq 1, \tag{17}
\]

where

\[
P_k(x_1, \ldots, x_k) = \sum_{j=1}^{k} (-1)^{j-1} (j-1)! B_{k,j}(x_1, \ldots, x_{k-j+1}) \tag{18}
\]

and \( \{B_{k,j}\} \) are the partial exponential Bell polynomials (Charalambides, 2002). Let us recall that, for a fixed positive integer \( k \) and \( j = 1, \ldots, k \), the \( j \)-th partial exponential Bell polynomial in the variables \( x_1, x_2, \ldots, x_{k-j+1} \) is a homogeneous polynomial of degree \( j \) given by

\[
B_{k,j}(x_1, \ldots, x_{k-j+1}) = \sum_{\lambda_1!\lambda_2! \cdots \lambda_{k-j+1}!} \frac{k!}{\lambda_1! \lambda_2! \cdots \lambda_{k-j+1}!} \prod_{i=1}^{k-j+1} \left( \frac{x_i}{i!} \right)^{\lambda_i} \tag{19}
\]

where the sum is taken over all sequences \( \lambda_1, \lambda_2, \ldots, \lambda_{k-j+1} \) of non negative integers such that

\[
\lambda_1 + 2\lambda_2 + \cdots + (k-j+1)\lambda_{k-j+1} = k, \quad \lambda_1 + \lambda_2 + \cdots + \lambda_{k-j+1} = j. \tag{20}
\]

The \( k \)-th logarithmic polynomial (18) is a special case of the \( k \)-th general partition polynomial

\[
G_k(a_1, \ldots, a_k; x_1, \ldots, x_k) = \sum_{j=1}^{k} a_j B_{k,j}(x_1, \ldots, x_{k-j+1}), \quad k \geq 1 \tag{21}
\]

when \( a_j = (-1)^{j-1} (j-1)! \) for \( j \geq 1 \). The first five general partition polynomials \( \{G_k\}_{k=1}^{5} \) are given in Table 1.

| \( k \) | \( G_k(a_1, \ldots, a_k; x_1, \ldots, x_k) \) |
|---|---|
| 1 | \( a_1 x_1 \) |
| 2 | \( a_1 x_2 + a_2 x_1^2 \) |
| 3 | \( a_1 x_3 + 3a_2 x_2 x_1 + a_3 x_1^3 \) |
| 4 | \( a_1 x_4 + 4a_2 x_3 x_1 + 6a_3 x_2 x_1^2 + a_4 x_1^4 + 3a_2 x_2^2 \) |
| 5 | \( a_1 x_5 + 5a_2 x_4 x_1 + 10a_2 x_3 x_2 + 10a_3 x_2 x_1^2 + 15a_3 x_1^3 + 10a_4 x_2 x_1^3 + a_5 x_1^5 \) |
If we set \( y_k = P_k(x_1, \ldots, x_k) \) for \( k \geq 1 \), then

\[
x_k = Y_k(y_1, \ldots, y_k) = \sum_{j=1}^{k} B_{k,j}(y_1, \ldots, y_{k-j+1}), \quad k \geq 1
\]

are the inverse relations, with \( \{ B_{k,j} \} \) given in Eq. (19). The polynomial \( Y_k \) is the \( k \)-th complete Bell (exponential) polynomial and is a special case of \( G_k \) in Eq. (21) when \( a_j = 1 \) for \( j \geq 1 \).

The logarithmic and the complete Bell polynomials allow us to deal with moments and cumulants of \( T \). Indeed if \( \tilde{g} \) is the Laplace transform of the PDF \( g(t) \) and the rhs of Eq. (15) is its Taylor expansion about 0, then \( \tilde{g}_0 = 1 \),

\[
\tilde{g}_k = (-1)^k \mathbb{E}[T^k], \quad k \geq 1
\]

and there exist cumulants of any order \( \{ c_k(T) \} \), see for instance Abate and Whitt (1996). In particular, from Eq. (16) and Eq. (17), we have

\[
c_k = (-1)^k P_k(\mathbb{E}[T], \ldots, \mathbb{E}[T^k]) = (-1)^k c_k(T), \quad k \geq 1.
\]

Vice-versa, if cumulants \( \{ c_k(T) \} \) are known, moments of \( T \) might be computed by using the inverse relations (22)

\[
\mathbb{E}[T^k] = Y_k(c_1[T], \ldots, c_k[T]), \quad k \geq 1
\]

or the well-known recursion formula (Di Nardo and Senato, 2006)

\[
\mathbb{E}[T^k] = c_k(T) + \sum_{i=1}^{k-1} \binom{k-1}{i-1} c_i(T) \mathbb{E}[T^{k-i}].
\]

If \( T \) is the FPT random variable of a Feller process modeled by Eq. (8), the following theorem gives the closed-form expression of the \( k \)-th cumulant for any order \( k \geq 1 \).

**Theorem 1.** The \( k \)-th FPT cumulant for the Feller process in Eq. (8) is

\[
c_k(T) = \left( -\frac{1}{\tau} \right)^k \left[ c_k^*(y_0) - c_k^*(S) \right]
\]

where

\[
c_k^*(w) = P_k \left[ h_1 \left( \frac{2 \tau (w - c)}{\sigma^2} \right), h_2 \left( \frac{2 \tau (w - c)}{\sigma^2} \right), \ldots, h_k \left( \frac{2 \tau (w - c)}{\sigma^2} \right) \right],
\]

\( P_k \) is the \( k \)-th logarithmic polynomial (18) and

\[
h_j(y) = j! \sum_{n \geq j} \binom{n}{j} \frac{y^n}{n!(\langle s \rangle_n)^j}, \quad j = 1, 2, \ldots, k,
\]

with \( \binom{n}{j} \) the unsigned Stirling numbers of first type and \( \langle s \rangle_n \) the \( n \)-th rising factorial.
Remark 1. Note that (Charalambides, 2002)

\[
\binom{n}{j} = B_{n,j}(0!, 1!, \ldots, (n-j+1)!), \ n \geq j
\]

with \(\{B_{n,j}\}\) the partial exponential Bell polynomials given in Eq. (19).

Proof. In Eq. (12), set \(A = 2\tau(y_0 - c)/\sigma^2\) and \(B = 2\tau(S - c)/\sigma^2\). From Eq. (24) we get

\[
\log \tilde{g}(z) = \log \frac{1}{1 + F_1\left(\frac{z}{\tau}; s; A\right)} = \sum_{k \geq 1} \left(-1\right)^k c_k(T) \frac{z^k}{k!},
\]

where

\[
1_F\left(\frac{z}{\tau}; s; y\right) = \sum_{n \geq 0} \frac{\langle z \rangle_n y^n}{\langle s \rangle_n n!}, \ y = A, B.
\]

To expand the rhs of Eq. (32) in formal power series in \(z\), observe that

\[
\langle \frac{z}{\tau} \rangle_n = \sum_{j=0}^{n} \binom{n}{j} \frac{z^j}{\tau^j}
\]

where \(\binom{n}{j}\) are the unsigned Stirling numbers of the first type. Replacing Eq. (33) in Eq. (32), after some algebra, we get

\[
1_F\left(\frac{z}{\tau}; s; y\right) = 1 + \sum_{k \geq 1} \frac{z^k}{\tau^k} \left( \sum_{n \geq k} \left[ \binom{n}{k} \frac{y^n}{\langle s \rangle_n n!} \right] \right).
\]

From Eqs. (16) and (17), we get

\[
\log 1_F\left(\frac{z}{\tau}; s; y\right) = \sum_{k \geq 1} \frac{P_k[h_1(y), \ldots, h_k(y)]}{k!} \frac{z^k}{\tau^k},
\]

where \(P_k\) is the \(k\)-th logarithmic polynomial given in Eq. (18) and \(h_j(y)\) is given in Eq. (29). Moreover Eq. (27) follows taking into account Eq. (31) and by observing that

\[
\log 1_F\left(\frac{z}{\tau}; s; A\right) = \log 1_F\left(\frac{z}{\tau}; s; A\right) - \log 1_F\left(\frac{z}{\tau}; s; B\right)
\]

\[
= \sum_{k \geq 1} \left( \frac{P_k[h_1(A), \ldots, h_k(A)] - P_k[h_1(B), \ldots, h_k(B)]}{\tau^k} \frac{z^k}{k!} \right).
\]

\[\square\]

Corollary 1. The mean FPT and the variance of \(T\) are respectively

\[
c_i(T) = \left(-\frac{1}{\tau}\right)^i \sum_{n \geq i} a_{i,n} \left(\frac{2\tau}{\sigma^2}\right)^n \left[(y_0 - c)^n - (S - c)^n\right], \ i = 1, 2
\]
where for \( n \geq i \)
\[
a_{i,n} = \begin{cases} 
\frac{n-1}{\langle s \rangle_n}, & i = 1 \\
2(n-1)H_{n-1} - \sum_{k=1}^{n-1} k^{-1} \frac{(n-k)^{-1}}{\langle s \rangle_k} \frac{1}{\langle s \rangle_{n-k}}, & i = 2 
\end{cases}
\] (37)

with \( H_{n-1} = \sum_{j=1}^{n-1} j^{-1} \) the harmonic numbers.

**Proof.** The mean FPT is obtained choosing \( k = 1 \) in Eq. (27) and observing that \( \binom{n}{1} = (n-1)! \) for \( n \geq 1 \). The variance of the FPT is obtained choosing \( k = 2 \) in Eq. (27), observing that \( \binom{n}{2} = (n-1)!H_{n-1} \) for \( n \geq 2 \) with \( H_n \) the \( n \)-th harmonic number and
\[
h_1(y)^2 = \sum_{n \geq 2} \left( \sum_{k=1}^{n-1} a_{1,k}a_{1,n-k} \right) y^n.
\] (38)

\[ \square \]

**Remark 2.** Observe that Eq. (36) gives \( \mathbb{E}[T] \) for \( i = 1 \) and coincides with the expression (14) of the first moment of \( T \). The comparison follows easily from the definition of \( s \) and the property \( \langle s \rangle_n = \Gamma(s+n)/\Gamma(s) \).

**Corollary 2.** If \( \{c_k(T)\} \) is the FPT cumulant sequence, then
\[
\mathbb{E}[T^k] = \frac{(-1)^k}{\tau^k} \sum_{i=0}^{k} \binom{k}{i} Y_{k-i}[c_1^*(y_0), \ldots, c_{k-i}^*(y_0)] Y_i[-c_1^*(S), \ldots, -c_i^*(S)]
\] (39)

where \( \{Y_i\} \) are the complete Bell polynomials given in Eq. (22) and \( Y_0 = 1 \).

**Proof.** Since \( Y_k(ay_1, a^2y_2, \ldots, a^ky_k) = a^kY_k(y_1, y_2, \ldots, y_k), a \in \mathbb{R} \), from Eq. (25) and Eq. (27), we get
\[
\mathbb{E}[T^k] = \frac{(-1)^k}{\tau^k} Y_k[c_1^*(y_0) - c_1^*(S), \ldots, c_k^*(y_0) - c_k^*(S)].
\] (40)

Applying the binomial type property of the complete Bell polynomials, we have
\[
Y_k[c_1^*(y_0) - c_1^*(S), c_2^*(y_0) - c_2^*(S), \ldots, c_k^*(y_0) - c_k^*(S)]
= \sum_{i=0}^{k} \binom{k}{i} Y_{k-i}[c_1^*(y_0), \ldots, c_{k-i}^*(y_0)] Y_i[-c_1^*(S), \ldots, -c_i^*(S)]
\] (41)

and the result follows. \[ \square \]

Note that
\[
Y_i[-y_1, \ldots, -y_i] = \sum_{j=1}^{i} (-1)^j B_{i,j}[y_1, \ldots, y_{i+j-1}], \ i \geq 1
\] (42)
since \( B_{i,j}(ay_1, \ldots, ay_{i+j-1}) = a^j B_{i,j}(y_1, \ldots, y_{i+j-1}), a \in \mathbb{R} \) from Eq. (19).
### 3.1. Computing FPT cumulants

For the subsequent applications of Theorem 1, we add some remarks on the efficiency of the implementation of Eq. (27). The logarithmic partition polynomials \( \{P_k\} \), with \( P_1(x_1) = x_1 \), might be generated by using the recurrence relation (Charalambides, 2002)

\[
P_k(x_1, \ldots, x_k) = x_k - \sum_{r=1}^{k-1} \binom{k-1}{r} x_r P_{k-r}(x_1, \ldots, x_{k-r}), \quad k \geq 2.
\]

(43)

About the computation of \( \{h_k(y)\} \) in Eq. (29), from Eq. (34), note that

\[
h_k(y) = \left. \frac{\partial^k}{\partial u^k} \, _1F_1(u; s; y) \right|_{u=0}, \quad k \geq 1.
\]

(44)

Derivatives of the Kummer’s function with respect to the parameter \( u \) have been computed in Ancarani and Gasaneo (2008). The special case \( u = 0 \) is given in terms of generalized Kampé de Fériet-like hypergeometric functions. An algorithm for the computation of the \( k \)-th derivative of the Kummer’s function is given in Abad and Sesma (2003). Here, we propose to use a standard implementation of the series in Eq. (29) involving the unsigned Stirling number of first type, as procedures implementing the well-known triangular recurrence relation (Charalambides, 2002)

\[
\left[ \begin{array}{c} n+1 \\ j \end{array} \right] = \left[ \begin{array}{c} n \\ j \end{array} \right] + n \left[ \begin{array}{c} n \\ j \end{array} \right], \quad j = 1, \ldots, n+1, \quad n \geq 0
\]

(45)

are available in many classical technical computing systems as Mathematica or R.

From Eq. (44) it turns out that the usefulness of expression (29) for the functions \( h_j \) is twofold. It also constitutes an alternative way to express the derivative in Eq. (44) and so it can simplify the form of the Kampé de Fériet function for particular values of the involved parameters. Moreover from Eq. (44) and the following expression (Abramowitz and Stegun, 1964)

\[
\frac{d^k}{dx^k} f(x) = k! \sum_{n \geq k} \frac{(-1)^{n-k} \binom{n}{k}}{n!} \Delta^n f(x)
\]

(46)

we infer the following formula for the forward differences of order \( n \)-th of the Kummer function:

\[
\Delta^n_u _1F_1(u; s; y) \bigg|_{u=0} = \frac{y^n}{(s)_n}.
\]

(47)

### 4. THE LAGUERRE-GAMMA POLYNOMIAL APPROXIMATION

The Edgeworth expansion is widely used in the literature to approximate a PDF around the Gaussian PDF, using a linear combination of Hermite polynomials with coefficients depending on the cumulants of the target PDF. To
approximate a non-Gaussian PDF, a different family of polynomials is necessary together with a different reference density (Asmussen, Goffard, and Laub, 2019). If the target PDF $g(t)$ is unknown but expected to be close to some reference density $\varphi(t)$, then $\varphi(t)$ is used as a first approximation to $g(t)$ and later the approximation is improved by using suitable correction terms depending on a set of orthonormal polynomials. The following theorems show how to approximate the FPT PDF of a Feller process by using as reference density the gamma PDF with scale parameter $\alpha + 1 > 0$ and shape parameter $\beta > 0$

$$\varphi_{\alpha,\beta}(t) = \frac{\beta^{\alpha+1}}{\Gamma(\alpha + 1)} t^\alpha e^{-\beta t}, \quad t > 0.$$  \hfill (48)

**Theorem 2.** Let $a_k^{(\alpha)} = E[Q_k^{(\alpha)}(\beta T)]$, $k \geq 0$ where

$$Q_k(t) = (-1)^k \left( \frac{\Gamma(\alpha + 1 + k)}{k! \Gamma(\alpha + 1)} \right)^{-1/2} L_k^{(\alpha)}(t),$$  \hfill (49)

and $L_k^{(\alpha)}(t)$ is the $k$-th generalized Laguerre polynomial

$$L_k^{(\alpha)}(t) = \Gamma(\alpha + 1 + k) \frac{k!}{k!} \sum_{j=0}^{k} \binom{k}{j} \frac{(-t)^j}{\Gamma(\alpha + j + 1)}.$$  \hfill (50)

For $t > 0$ the series

$$U(\beta t, r) := \sum_{k \geq 0} a_k^{(\alpha)} Q_k^{(\alpha)}(\beta t) r^k$$  \hfill (51)

converges if $r \in (0, 1)$ and

$$\lim_{r \to 1} U(\beta t, r) = \frac{g(t)}{\varphi_{\alpha,\beta}(t)}.$$  \hfill (52)

**Proof.** Set $\beta t = w$ and observe that in Eq. (51) the series might be rewritten as

$$U(w, r) = \sum_{k \geq 0} b_k^{(\alpha)} L_k^{(\alpha)}(w) r^k, \quad w > 0$$  \hfill (53)

where for $k \geq 0$

$$b_k^{(\alpha)} = \frac{\Gamma(k+1)}{\Gamma(k+1+\alpha)} \int_0^\infty e^{-t} t^\alpha L_k^{(\alpha)}(t) f(t) \, dt$$  \hfill (54)

with $f(t) = g_\beta(t)/\varphi_{\alpha,1}(t), g_\beta(t)$ the PDF of $\beta T$ and $\varphi_{\alpha,1}(t)$ as given in Eq. (48). A sufficient condition to have the convergence of the series (53) for $r \in (0, 1)$ at every point of continuity of $f(t)$ is (Hille, 1926)

$$\int_0^\infty e^{-zt} t^\alpha f(t) \, dt = \Gamma(\alpha + 1) \int_0^\infty e^{-(z-1)t} g_\beta(t) \, dt < \infty$$  \hfill (55)
for every $z - 1 > 0$, which is fulfilled when $T$ is the FPT random variable of a Feller process from Eq. (12). Therefore
\[
\lim_{r \to 1} U(w, r) = \frac{g_\beta(w)}{\varphi_{\alpha,1}(w)} \tag{56}
\]
and Eq. (52) follows from Eq. (56) after some algebra, replacing $w$ by $\beta t$ and recalling that $\beta g_\beta(\beta t) = g(t)$.

Eqs. (51) and (52) justify the approximation of $g$ with the polynomial of degree $n$
\[
\hat{g}(t) := \varphi_{\alpha,\beta}(t) \sum_{k=0}^{n} a_k^{(\alpha)} Q_k^{(\alpha)}(\beta t) \tag{57}
\]
for a suitable choice of $n$, that we discuss in the next section.

**Remark 3.** The polynomial approximation (57) is particularly suited when the PDF of $T$ is unknown, but its moments are available, as happens for the FPT random variable of the Feller process thanks to Corollary 2. Indeed by observing that
\[
\mathbb{E}[Q_k^{(\alpha)}(\beta T)] = (-1)^k \left( \frac{\Gamma(\alpha + 1 + k)}{k! \Gamma(\alpha + 1)} \right)^{1/2} \int_0^\infty g(t) L_k^{(\alpha)}(\beta t) dt
\]
\[
= (-1)^k \left( \frac{\Gamma(\alpha + 1) \Gamma(\alpha + 1 + k)}{k!} \right)^{1/2} \sum_{j=0}^{k} \frac{k}{j} \left( -\beta \right)^j \mathbb{E}(T^j) \Gamma(\alpha + j + 1),
\]
some algebra allows us to rewrite $\hat{g}(t)$ in Eq. (57) as
\[
\hat{g}(t) = \beta(\beta t)^\alpha e^{-\beta t} \sum_{k=0}^{n} A_k^{(\alpha)} L_k^{(\alpha)}(\beta t), \ t > 0 \tag{58}
\]
with coefficients
\[
A_k^{(\alpha)} = \sum_{j=0}^{k} \binom{k}{j} \left( -\beta \right)^j \mathbb{E}(T^j) \Gamma(\alpha + j + 1), \ k = 0, 1, \ldots, n \tag{59}
\]
depending on the moments of $T$. Note that $\{A_k^{(\alpha)}\}$ might be expressed directly in terms of cumulants of $T$ by using Eq. (25).

Sufficient conditions for the convergence of the series
\[
\sum_{k=0}^{n} a_k^{(\alpha)} Q_k^{(\alpha)}(\beta t), \ t > 0 \tag{60}
\]
can be recovered by using the analogous on the Laguerre series (Hille, 1926). Indeed in such a case we have $\lim_{r \to 1} U(\beta t, r) = U(\beta t, 1)$ and
\[
g(t) = \varphi_{\alpha,\beta}(t) \sum_{k=0}^{n} a_k^{(\alpha)} Q_k^{(\alpha)}(\beta t), \ t > 0. \tag{61}
\]
The next corollary gives a sufficient condition on $g(t)$ to have the series representation (61).
Corollary 3. The PDF \( g(t) \) has the series representation (61) if

\[
\int_0^\infty t^{-\alpha} e^{\beta t} g(t)^2 dt < \infty.
\]  

(62)

Proof. Condition (62) is equivalent to ask \( g(t) / \varphi_{\alpha,\beta}(t) \in \mathcal{L}^2(\nu) \), equipped with the usual inner product \( \langle g_1, g_2 \rangle = \int g_1 g_2 d\nu \) and \( \nu \) the measure having density \( \varphi_{\alpha,\beta}(t) \). As \( \nu \) admits moment generating function and all its moments are finite, there exists a complete set of orthonormal polynomials in \( \mathcal{L}^2(\nu) \), such that if \( g / \varphi_{\alpha,\beta} \in \mathcal{L}^2(\nu) \), we may expand \( g(t) / \varphi_{\alpha,\beta}(t) \) in terms of these polynomials. Let us observe that \( \{ Q_{k}^{(\alpha)}(\beta) \} \) is a family of orthonormal polynomials in \( \mathcal{L}^2(\nu) \) since \( \{ L_k^{(\alpha)}(t) \} \) is a family of orthogonal polynomials with respect to the weight function \( t^\alpha e^{-t} \). Therefore the Laguerre series (60) with \( a_k^{(\alpha)} = \mathbb{E}[Q_k^{(\alpha)}(\beta T)] \), \( k \geq 0 \) represents the Fourier-Laguerre expansion of \( g(t) / \varphi_{\alpha,\beta}(t) \) from whose uniqueness Eq. (61) follows. \( \square \)

As \( g(t) \) is unknown, it’s not easy to verify directly the condition (62). If

\[
\sum_{k=0}^{\infty} \left| \mathbb{E}[Q_k^{(\alpha)}(\beta T)] \right| < \infty
\]

(63)

then expansion (61) holds, due to the Parseval identity. The accuracy of the approximation (57) depends upon the decay rate of \( \{ \mathbb{E}[Q_k^{(\alpha)}(\beta T)] \} \), as the \( \mathcal{L}^2(\nu) \)-loss is \( \sum_{k=n+1}^{\infty} (\mathbb{E}[Q_k^{(\alpha)}(\beta T)])^2 \) for a given order of truncation \( n \). A sufficient condition to have Eq. (62) is

\[
(\mathbb{E}[Q_k^{(\alpha)}(\beta T)])^2 \approx k^{(-1-\varepsilon)} \text{ as } k \to \infty \text{ and } \varepsilon > 0.
\]

(64)

So it is fundamental to have a good algorithm to evaluate the coefficients \( \{ \mathbb{E}[Q_k^{(\alpha)}(\beta T)] \} \). This issue will be analyzed in the next paragraph.

4.1. Computational issues

To simplify the implementation, the approximating polynomial \( \hat{g}(t) \) has been computed by using Eq. (58). The first five generalized Laguerre polynomials are given in Table 2. Many packages\(^1\) return the first \( n \) generalized La-

| Table 2. Generalized Laguerre polynomials |
|-----|----------------|
| \( k \) | \( L_k^{(\alpha)}(t) \) |
| 1 | \( (\alpha + 1)1 - t \) |
| 2 | \( ((\alpha + 1)2 - 2(\alpha + 2)1t + t^2) / 2! \) |
| 3 | \( (\alpha + 1)3 - 3(\alpha + 2)2t + 3(\alpha + 3)1t^2 - t^3) / 3! \) |
| 4 | \( (\alpha + 1)4 - 4(\alpha + 2)3t + 6(\alpha + 3)2t^2 - 4(\alpha + 4)1t^3 + t^4) / 4! \) |
| 5 | \( (\alpha + 1)5 - 5(\alpha + 2)4t + 10(\alpha + 3)3t^2 - 10(\alpha + 4)2t^3 + 5(\alpha + 5)1t^4 - t^5) / 5! \) |
guerre polynomials by using the following recursion formula (Charalambides, 2002)

\[ L_{k+1}^{(\alpha)}(t) = \frac{(2k + 1 + \alpha - t)L_k^{(\alpha)}(t) - (k + \alpha)k!}{k + 1}, \quad k \geq 0, \quad (65) \]

with \( L_0^{(\alpha)}(t) = 1 \). The same recursion (65) allows an efficient computation of the coefficients \( \{A_k^{(\alpha)}\} \). This result is proved in the following lemma where we use the symbolic calculus (Charalambides, 2002) formalized through the employment of a linear operator acting on a ring of polynomials, for details see Di Nardo (2012).

**Proposition 1.** Let \( A_k^{(\alpha)}(y) = \frac{k!}{\Gamma(\alpha + 1 + k)}L_k^{(\alpha)}(y) \), for \( k \geq 0 \).

Then

\[ A_k^{(\alpha)} = E[A_k^{(\alpha)}(\beta m)] \quad (66) \]

where \( E \) is a linear operator transforming \( m^j \) in \( m_j = E[T^j] \), that is \( E[m^j] = E[T^j], j \geq 1 \) and \( E[1] = m_0 = 1 \).

**Proof.** The result follows from Eq. (59), since

\[ E[A_k(\beta m)] = \sum_{j=0}^{k} \binom{k}{j} (-\beta)^j E[m^j] \frac{1}{\Gamma(\alpha + j + 1)} \quad (67) \]

and using the linear operator \( E \).

Note that the moments \( \{m_j\} \) are calculated from cumulants using the recursive relation (26).

The question of how to select the parameters \( \alpha \) and \( \beta \) in Eq. (58) results to be a crucial point. A general guideline to their selection consists in matching the first two moments of \( g(t) \) with the first two of \( \varphi_{\alpha,\beta}(t) \). From a statistical point of view this choice mimics the well-known method of moments. From a computational point of view, if the first two moments of \( g(t) \) and \( \varphi_{\alpha,\beta}(t) \) coincide, then \( A_k^{(\alpha)} = A_k^{(\alpha)} = 0 \) for \( k = 1, 2 \) simplifying the computation of Eq. (58) (Asmussen, Goffard, and Laub (2019)). According to this rule, if

\[ \beta := \frac{c_1(T)}{c_2(T)} \quad \text{and} \quad \alpha := \beta E[T] - 1 = \frac{c_2^2(T)}{c_2(T)} - 1 \quad (68) \]

then \( A_1^{(\alpha)} = A_1^{(\alpha)} = 0 \) in Eq. (58).

The choice of \( \alpha \) and \( \beta \) in Eq. (68) deserves some deeper analysis. First note that the shape parameter \( \alpha + 1 \) is given by the mean of the scaled random variable \( \beta T \). Instead, the rate parameter \( \beta \) measures the inverse relative variance of \( g(t) \). The relative variance of a PDF is a normalized measure of its dispersion. Thus, the more spread out is \( \varphi_{\alpha,\beta}(t) \) the greater is the underdispersion of \( g(t) \). The next section gives some examples and applications of Eq. 1 See for example the package orthopolynom in R.
(58) stopped at \( n = 5 \). The motivation of this choice stems from the statistical meaning of the coefficients \( \{A_k^{(\alpha)}\} \). As \( \alpha := \mathbb{E}[\beta T - 1] \), symbolic calculus shows that the coefficients \( \{A_k^{(\alpha)}\} \) are related to the \( k \)-th moment \( \mathbb{E}[(\beta T - 1)^k] \), without the normalizing constant \( \Gamma(\alpha + j + 1) \), which depends on the orthonormal property of \( \{Q_k^{(\alpha)}(t)\} \). Thus the third-order coefficient \( A_3^{(\alpha)} \) accounts for the skewness of \( g(t) \) while the fourth-order coefficient \( A_4^{(\alpha)} \) involves the weight of tails in causing dispersion, that is the kurtosis. The fifth-order coefficient \( A_5^{(\alpha)} \) involves the hyper-skewness \( m_5 \) of \( g(t) \) (Khademaloomoom, Narayan, and Sharma, 2019). Hyper-skewness measures the asymmetric sensitivity of the kurtosis, that is the relative importance of tails versus the center in causing skewness. Note that the sixth moment \( m_6 \) in \( A_6^{(\alpha)} \) is the PDF hyper-kurtosis and measures both the peakedness and the tails compared with the normal distribution. As we are considering PDFs with support \((0, \infty)\), the contribution of this coefficient is not statistically meaningful and not considered here.

Some suitable choices of the rate parameter \( \beta \) might improve the approximation, as the following propositions show.

**Proposition 2.** If \( g(\varphi_{\alpha,\beta}) \in \mathcal{L}^2(\nu) \), then \( \beta < \frac{2}{\mathbb{E}[T]} \).

**Proof.** As \( g(\varphi_{\alpha,\beta}) \in \mathcal{L}^2(\nu) \), all the integrals in
\[
\int_0^{\infty} t^{-\alpha} e^{\beta t} g(t)^2 \, dt = \int_0^1 t^{-\alpha} e^{\beta t} g(t)^2 \, dt + \int_1^{\infty} t^{-\alpha} e^{\beta t} g(t)^2 \, dt = I_1 + I_2 \quad (69)
\]
are finite. The FPT PDF \( g(t) \) has exponential long-time behavior with parameter the inverse mean FPT (Masoliver and PerellÃ§s, 2014), i.e
\[
g(t) \approx \frac{1}{\mathbb{E}(T)} e^{-\frac{\beta t}{\mathbb{E}(T)}}. \quad (70)
\]
Therefore to have the convergence of the latter integral in Eq. (69), it is necessary to have \( \beta - \frac{2}{\mathbb{E}[T]} < 0 \). \( \square \)

Thus in the following proposition we give a sufficient condition for the ratio \( g(\varphi_{\alpha,\beta}) \) to be in \( \mathcal{L}^2(\nu) \), assuming
\[
\beta < \frac{2}{\mathbb{E}[T]} \iff 2c_2(T) > [c_1(T)]^2 \quad (71)
\]
under the choice (68).

**Proposition 3.** We have \( g(\varphi_{\alpha,\beta}) \in \mathcal{L}^2(\nu) \) if \( \beta < 2/\mathbb{E}[T] \) and \( g(t) = o(t^\delta) \) with \( 2\delta + 1 > |c_1(T)|^2/c_2(T) \).

**Proof.** Let us consider again Eq. (69). Under condition (71), the integral \( I_2 \) is always finite. Moreover \( 2c_2(T) > [c_1(T)]^2 \iff -1 < \alpha < 1 \). Let us focus on \( I_1 \). If \( \alpha \in (0, 1) \) and \( g^2(t) = o(t^{2\delta}) \) with \( 2\delta > \alpha \) then the integrand function in \( I_1 \) is limited and \( I_1 \) is finite. \( \square \)
5. EXAMPLES

As mentioned before, the Feller process plays a key role in a variety of applications. In this section we will investigate three examples coming from different areas of study.

5.1. Example 1

In the first example we consider dimensionless quantities to show the way the approximation is implemented and how it performs.

In Figure 5.1-top we show the PDF of the first passage time through $S = 1$ for the Feller process solution of Eq. (8) for $y_0 = 0.2$, $c = 0$, $\tau = 1/1.5$, $\sigma = 1$ and $\mu = 0.9$. The curves are obtained using just 2, 3, 4 or 5 cumulants in the approximation method, i.e. from Eq. (58) for $n = 2, 3, 4, 5$. The approximations are compared to the FPT PDF obtained through simulation of $10^4$ first passage times of the process by discretization of Eq. (8) (see the Appendix). We observe that the agreement is satisfactory even for small $n$, although the expression of the cumulants of any order is available and in principle can be used to improve the approximation. The absolute error between the simulated PDF and the approximated one is shown in Figure 5.1-bottom. We observe that the error remains smaller than 0.05 for $t > 0$. In $t = 0$ the error is bigger, but the difference is due mainly to the error in the simulation rather than in the approximation. The reason is in the criteria for the bandwidth chosen in building the density from the histograms of the simulated first passage times.

5.2. Example 2: Neuronal modeling

The Feller model was proved to fit experimental data of in vitro neurons under different conditions (Hopfner, 2007). For this reason in this subsection we focus on an application to neuronal modeling of Eq. (8). The solution process $Y_t$ describes the evolution in time of the depolarization of the membrane potential of the neuron that is modelled as a leaky RC circuit with a drift characterizing the input stimuli. Eq. (8) describes the membrane depolarization until the occurrence of a spike. In accordance with the model, the spikes are generated when the process $Y_t$ crosses a voltage threshold $S$ for the first time, involving thus the FPT random variable. The process is reset to the starting point $y_0$ after the spike and the evolution starts anew. In this framework $y_0$ is the starting depolarization, $\sigma$ determines the amplitude of the noise, $c$ is the inhibitory reversal potential, $\tau$ is the inverse of the characteristic time constant of the neuron that takes into account the spontaneous voltage decay towards the resting potential in the absence of inputs and $\mu$ characterizes the input the neuron under consideration receives. In the following, we consider the same parameters values used in Lansky, Sacerdote, and Tomassetti (1995), the resetting potential is equal to zero, i.e. $y_0 = 0$ mV, the inhibitory reversal potential is fixed to $c = -10$ mV, the noise amplitude $\sigma = 1.2 \sqrt{\text{mV}/\sqrt{\text{ms}}}$,
Figure 1. TOP: Density of the first passage time through $S = 1$ for the Feller process solution of Eq. (8) for $y_0 = 0.2$, $c = 0$, $\tau = 1/1.5$, $\sigma = 1$ and $\mu = 0.9$. The curves are obtained from Eq. (58) for $n = 2, 3, 4, 5$ (in the legend). The approximations are compared to the simulated FPT PDF (in solid-red). The curve in red is built by simulation of $10^4$ first passage times of the Feller process obtained discretizing Eq. (8) by means of Eq. (A2) with $\Delta t = 10^{-2}$. BOTTOM: The absolute error between the simulated PDF and the approximated one for the above cases.

$\mu = 3$ mV/ms and the firing threshold to $S = 10$ mV. The parameter of spontaneous decay is chosen $\tau = 0.2$ ms (Figure 5.2).

In Ditlevsen and Lansky (2006) the noise amplitude is chosen $\sigma = 2 \sqrt{\text{mV}/\text{ms}}$ (Figure 5.3). In this case the density is more skewed and the approximation fails to fit well the mode, although the error remains of the order of 0.05 (not shown). The reason is in the properties of the gamma distribution that is our reference distribution: the mode is Indeed not defined for $\alpha < 1$, ($\alpha = 0.07$ in the example).

Another property of the gamma distribution to take care of is the shape of the distribution for small $\alpha$. In fact in this case the gamma density stops to have the typical bell-shape, and $\varphi_{\alpha,\beta}(t)$ might fail a good approximation of $g(t)$ for small $t$. This situation is presented in the following example.
Figure 2. Density of the first passage time through $S = 10$ mV for the Feller neuronal model (8) for $y_0 = 0$ mV, $c = -10$ mV, $\tau = 0.2$ ms, $\sigma = 1.2 \sqrt{\text{mV}/\sqrt{\text{ms}}}$ and $\mu = 3$ mV/ms (the regime is suprathreshold). The curves are obtained from Eq. (58) for $n = 2, 3, 4, 5$ (in the legend). The approximations are compared to the simulated FPT PDF (in solid-red).

Figure 3. Density of the first passage time for the Feller neuronal model (8) for the same choice of parameters of Figure 5.2 except for $\sigma = 2 \sqrt{\text{mV}/\sqrt{\text{ms}}}$ and $\mu = 4$ mV/ms. The discretization step for the simulation is $\Delta t = 10^{-3}$.

5.3. Example 3: Financial mean-reverting models

In mathematical finance the Feller process goes under the name of CIR model and it is used to study the term structure of interest rates (Cox, Ingersoll, and Ross, 1985) or mean-reverting models for a credit spread (Linetsky, 2004). In the option pricing literature, the Feller process is used to describe the variance in models with stochastic volatility, where the most notable example is probably the Heston model (Heston, 1993; Rouah, 2013). In this example we consider a stochastic model for an instantaneous credit spread following Eq. (8) in $t \in [0.01, 4]$ with the long-run credit spread level of 200 basis points ($\mu = 0.02 \cdot 0.25$), the initial spread level of 100bp ($y_0 = 0.01$), the rate of mean-reversion $\tau = 0.25$, and the volatility parameter $\sigma = 0.1$ (parameters given in
Linetsky (2004)). We are interested in the first passage time density of the long-run level \( S = 0.02 \), starting from \( y_0 < S \). Comparing the plot in Figure 5.4 with the one given by Linetsky (2004) obtained with a different method of approximation, we observe a good asymptotic agreement. We stress that the term asymptotic can be misleading since the agreement is good already for relatively small \( t \). The PDF shape is not preserved for \( t \approx 0.01 \). The reason is that for \( \alpha < 1 \) (\( \alpha = -0.34 \) in this case) the mode of the gamma distribution is not defined, and thus it cannot be matched with the one of \( g(t) \), if it exists. However if we impose the theoretical information that \( g(0) = 0 \), the behaviour of the PDF is reproduced. Note that using Eq. (58) we overcome the difficulties arisen from the simulation and the need to use 52 terms in approximation expansion suggested by Linetsky.

![Figure 4. Density of the FPT for the CIR model (8) in the interval \( t \in [0.01, 4] \) as in Linetsky (2004), i.e. \( \mu = 0.02 \cdot 0.25, y_0 = 0.01, \tau = 0.25, \sigma = 0.1 \) and \( S = 0.02 \). The curve is obtained from Eq. (58) for \( n = 5 \).](image)

6. CONCLUSIONS AND OPEN PROBLEMS

We considered the well-known Feller stochastic process and the related FPT problem through a constant boundary. We provided a manageable closed form expression for the cumulants of \( T \) of any order by which moments can be easily obtained, improving the current results available for the first three moments only. Note that the knowledge of higher moments gives qualitative information on the FPT PDF such as skewness, kurtosis, hyper-skewness and hyper-kurtosis.

We used cumulants to build a polynomial approximation of the FPT PDF \( g(t) \), whose expression in closed form is still missing in the literature. The method is carried out involving the gamma distribution as a first approximation to \( g(t) \) and then improving this approximation by adding suitable correction terms based on a set of Laguerre polynomials. The resulting Laguerre-Gamma polynomial has coefficients whose computation was lightened by using the well-known recurrence relation of the Laguerre polynomials and the
symbolic calculus. This computation was further simplified choosing the parameters of the gamma distribution with the method of moments. We have shown that the proposed method allows us to obtain good approximation of \( g(t) \) even using a low degree (5 in the analyzed case-studies). Moreover it overcomes the difficulties arisen from the simulation for time \( t \) close to zero. Some care must be taken when the PDF \( g(t) \) is expected to have a mode differently from the gamma distribution selected from the choice of its parameters. This circumstance deserves to be further investigated either in the choice of the parameters and in the expected properties of \( g(t) \). Moreover, we give sufficient conditions to improve the approximation of the PDF \( g(t) \) with the Laguerre-Gamma polynomial; criteria that are fulfilled in most cases of application.

Future work includes the extension of this approach to other processes belonging to the class of Pearson’s diffusion, since the expression of the Laplace transform of the FPT PDF for these processes is often written as a ratio of two hypergeometric functions. More in general, when the Laplace transform of the FPT PDF is a ratio of functions admitting a power series representation, cumulants might be recovered by using the algebra of formal power series and different polynomial approximations might be tested. For example if the transition PDF of the process has a power series representation of the Laplace transform \( \tilde{f}(z; x, y_0) = \int_0^\infty e^{-zt} f(x, t|y_0, 0)\, dt \), thus the proposed method might be investigated as \( \tilde{g}(z) \) is again a ratio of power series, that is \( \tilde{g}(z) = \tilde{f}(z; x, y_0)/\tilde{f}(z, x, S) \).

**Appendix A: The Milstein method**

To estimate the FPT PDF \( g(t) \), we have implemented a classical Monte Carlo method and simulated the paths of \( Y_t \) using the stochastic differential equation (8). The algorithm we refer relies on the Milstein scheme of discretization that is often used when the term \( A_2 \) of the SDE \( dY_t = A_1(Y_t, t)\, dt + \sqrt{A_2(Y_t, t)}\, dW(t) \) depends on the process \( Y_t \) (see for instance Kloeden and Platen (2011)). Truncation of the Itô-Taylor expansion at the second order produces Milstein’s method:

\[
Y_n = Y_{n-1} + A_1(Y_{n-1}) \Delta t + \sqrt{A_2(Y_{n-1})} \Delta W_{n-1} \nonumber
\]

\[
+ \frac{1}{2} \sqrt{A_2(Y_{n-1})} (\sqrt{A_2(Y_{n-1})})' [(\Delta W_{n-1})^2 - \Delta t] \quad \text{(A1)}
\]

for \( n = 1, 2, \ldots, N \) for some \( N \). The Milstein scheme exhibits convergence of order 1 in the strong sense and is a generalization of the Euler-Maruyama discretization scheme (the two methods coincide when \( A_2(Y_t) \) does not depend on \( Y_t \)). In case of Eq. (8), Eq. (A1) gives

\[
Y_n = Y_{n-1} + (\tau Y_{n-1} + \mu) \Delta t + \sigma \sqrt{Y_{n-1} - c} \, \Delta W_{n-1} + \frac{1}{4} \sigma^2 [(\Delta W_{n-1})^2 - \Delta t] \nonumber\]

\[
\quad \text{(A2)}
\]
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REFERENCES

Abad, J. and Sesma, J., “Successive derivatives of Whittaker functions with respect to the first parameter,” Computer Physics Communications 156, 13 – 21 (2003).

Abate, J. and Whitt, W., “An operational calculus for probability distributions via Laplace transforms,” Advances in Applied Probability 28, 75–113 (1996).

Abramowitz, M. and Stegun, I., Handbook of Mathematical Functions: With Formulas, Graphs, and Mathematical Tables (Dover Publications, 1964).

Ancarani, L. U. and Gasaneo, G., “Derivatives of any order of the confluent hypergeometric function $F_1(a, b, z)$ with respect to the parameter $a$ or $b$,” Journal of Mathematical Physics 49, 063508, 16 (2008).

Antunes, P., Ferreira, S. S., Ferreira, D., Nunes, C., and Mexia, J. T., “Estimation in additive models and ANOVA-like applications,” Journal of Applied Statistics 0, 1–10 (2020).

Arnold, L., Stochastic Differential Equations: Theory and Applications (Dover Publications, 2013).

Asmussen, S., Goffard, P.-O., and Laub, P. J., “Orthonormal polynomial expansions and lognormal sum densities,” in Risk and Stochastics (2019) Chapter 6, pp. 127–150, https://www.worldscientific.com/doi/pdf/10.1142/9781786341952_0008.

Charalambides, C. A., Enumerative Combinatorics (Chapman & Hall/CRC, Boca Raton, FL, 2002) pp. 609.

Cox, J. C., Ingersoll, J. E., and Ross, S. A., “A theory of the term structure of interest rates,” Econometrica 53, 385–407 (1985).

Di Nardo, E., “Symbolic calculus in mathematical statistics: a review,” Seminaire Lotharingien de Combinatoire 67, Art. B67a, 72 (2012).

Di Nardo, E. and Senato, D., “An umbral setting for cumulants and factorial moments,” European Journal of Combinatorics 27, 394 – 413 (2006).

Ditlevsen, S. and Lansky, P., “Estimation of the input parameters in the Feller neuronal model,” Physical Review E 73, 061910 (2006).

D’Onofrio, G., Lansky, P., and Pirozzi, E., “On two diffusion neuronal models with multiplicative noise: The mean first-passage time properties,” Chaos: An Interdisciplinary Journal of Nonlinear Science 28, 043103 (2018), https://doi.org/10.1063/1.5009574.

Feller, W., “Two Singular Diffusion Problems,” Annals of Mathematics 54, 173–182 (1951).

Forman, J. and Sørensen, M., “The Pearson diffusions: A class of statistically tractable diffusion processes,” Scandinavian Journal of Statistics 35, 438–465 (2008).

Giorno, V., Lansky, P., Nobile, A., and Ricciardi, L., “Diffusion approximation and first-passage-time problem for a model neuron,” Biological Cybernetics 58, 387–404 (1988).

Giorno, V., Nobile, A. G., Ricciardi, L. M., and Sacerdote, L., “Some remarks on the Rayleigh process,” Journal of Applied Probability 23, 398–408 (1986).
Going-Jaeschke, A. and Yor, M., “A survey and some generalizations of Bessel processes,” Bernoulli 9, 313–349 (2003).
Heston, S.L., “A closed-form solution for options with stochastic volatility with applications to bond and currency options,” The Review of Financial Studies 6(2), 327–343 (1993).
Hille, E., “On Laguerre’s series,” Proceedings of the National Academy of Sciences 12, 265–269 (1926), https://www.pnas.org/content/12/4/265.full.pdf.
Hopfner, R., “On a set of data for the membrane potential in a neuron,” Mathematical Biosciences 207, 275 – 301 (2007).
Karlin, S. and Taylor, H., A Second Course in Stochastic Processes, v. 2 (Academic Press, 1981).
Kent, J. T., “Eigenvalue expansions for diffusion hitting times,” Zeitschrift für Wahrscheinlichkeitstheorie und Verwandte Gebiete 52, 309–319 (1980).
Kent, J. T., “The spectral decomposition of a diffusion hitting time,” Annals of Probability 10, 207–219 (1982).
Khademalomoom, S., Narayan, P. K., and Sharma, S. S., “Higher moments and exchange rate behavior,” Financial Review 54, 201–229 (2019), https://onlinelibrary.wiley.com/doi/pdf/10.1111/fire.12171.
Kloeden, P. and Platen, E., Numerical Solution of Stochastic Differential Equations, Stochastic Modelling and Applied Probability (Springer Berlin Heidelberg, 2011).
Lansky, P., Sacerdote, L., and Tomassetti, F., “On the comparison of Feller and Ornstein-Uhlenbeck models for neural activity,” Biological Cybernetics 73, 457–465 (1995).
Linetsky, V., “Computing hitting time densities for CIR and OU diffusions: Applications to mean-reverting models,” Journal of Computational Finance 7, 1–22 (2004).
Martin, E., Behn, U., and Germano, G., “First-passage and first-exit times of a Bessel-like stochastic process,” Physical Review E 83, 051115 (2011).
Masoliver, J. and Perellàs, J., “First-passage and escape problems in the Feller process,” Physical Review E 86(4-1), 041116 (2014).
McCullagh, P., Tensor Methods in Statistics, (Chapman & Hall, London, 1987).
Øksendal, B., Stochastic Differential Equations: An Introduction with Applications (Springer, 1998).
Ramos-Alarcón, F. and Kontorovich, V., “First-passage time statistics of Markov gamma processes,” Journal of the Franklin Institute 350, 1686–1696 (2013).
Rouah, F.D., The Heston Model and Its Extensions in Matlab and C, (John Wiley & Sons, 2013).
Seneta, E., “Fitting the variance-gamma model to financial data,” Journal of Applied Probability 41, 177–187 (2004).
Siebert, A. J. F., “On the first passage time probability problem,” Physical Review 81, 617–623 (1951).