Anxiety and clinical outcomes of patients with acute coronary syndrome: a meta-analysis

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ABSTRACT

Objectives Anxiety has been suggested to be associated with poor outcomes in patients with ACS. However, results of previous follow-up studies were inconsistent. The aim of this meta-analysis was to evaluate the association between anxiety and clinical outcomes in patients with ACS, and to investigate the potential role of depression underlying the above association.

Design A meta-analysis of prospective follow-up studies.

Setting Hospitals.

Participants Patients with ACS.

Interventions We included related prospective follow-up studies up to 20 July 2019 that were identified by searching PubMed and Embase databases. A random-effect model was used for the meta-analysis.

Primary and secondary outcome measures We determined the association between anxiety and risks of mortality and adverse cardiovascular events (MACEs) in patients with ACS.

Results Our analysis included 17 studies involving 39038 patients with ACS. Anxiety was independently associated with increased mortality risk (adjusted risk ratio (RR) 1.21, 95% CI 1.07 to 1.37, p=0.002) and MACEs (adjusted RR 1.47, 95% CI 1.24 to 1.74, p<0.001) in patients with ACS. Subgroup analyses showed that depression may at least partly confound the association between anxiety and poor outcomes in patients with ACS. Adjustment of depression significantly attenuated the association between anxiety and MACEs (adjusted RR 1.25, 95% CI 1.04 to 1.52, p=0.02). Moreover, anxiety was not significantly associated with mortality risk after adjusting for depression (adjusted RR 0.88, 95% CI 0.66 to 1.17, p=0.37).

Conclusions Anxiety is associated with increased risk of mortality and MACEs in patients with ACS. However, at least part of the association may be confounded by concurrent depressive symptoms in these patients.

INTRODUCTION

Despite significant improvement in the management of patients with coronary artery disease (CAD) in recent decades, CAD, particularly acute coronary syndrome (ACS), remains one of the leading causes of death worldwide.1–3 ACS, including acute myocardial infarction (MI) and unstable angina (UA), refers to a category of severe CAD characterised by acute rupture of unstable atherosclerotic plaques and subsequent obstruction of the coronary artery lumen.4 Patients with ACS not only suffer from severe clinical symptoms of chest pain, but are also vulnerable to fatal complications such as malignant arrhythmia and haemodynamic instability.4 Moreover, invasive procedures, primarily including percutaneous coronary intervention or coronary artery bypass graft, are frequently performed in these patients.5 All of the above factors may contribute to the pathogenesis of affective disorders in these patients, such as depression and anxiety.6

Previous studies have confirmed that depression independently predicts poor clinical outcomes in patients with CAD.7 However, the influence of anxiety on prognosis of CAD patients, particularly for those with ACS, is poorly understood.8 Although the potential prognostic efficacy of anxiety for patients with ACS has been previously evaluated, results of these studies are inconsistent.9–25 Some studies indicated that anxiety is a risk factor of poor prognosis in patients with ACS,9 11 12 20 23 25 while others did not support this finding.10 13–19 21 22 24
Although two previous meta-analyses found that patients with ACS with anxiety may have a higher risk of mortality and other adverse outcomes compared with those without anxiety,26 27 these conclusions were mainly based on studies with univariate analyses, and one of the meta-analyses included a high risk of publication bias.26 However, many prospective studies with multivariable analyses have been performed to evaluate the effect of anxiety on prognosis in patients with ACS,16–25 providing rationale to perform an updated meta-analysis. Considering that anxiety and depression are highly correlated psychological disorders,28 it is important to determine the extent to which the association between anxiety and prognosis of patients with ACS is independent of depression. Accordingly, we performed an updated meta-analysis to evaluate the potential prognostic influence of anxiety on adverse clinical outcomes in patients with ACS. Moreover, we aimed to explore whether concurrent depression confounds the association between anxiety and adverse outcomes in patients with ACS.

METHODS
This study was designed as a meta-analysis of prospective observational studies, and was performed in accordance with the Meta-analysis of Observational Studies in Epidemiology29 and Cochrane’s Handbook30 guidelines.

Literature search
PubMed and Embase were searched for relevant records with the combination of the following terms: (1) “anxiety” OR “tension” OR “post-traumatic stress disorder” OR “panic” OR “phobia” OR “phobic” OR “worry”; (2) “myocardial infarction” OR “acute coronary syndrome” OR ACS OR “unstable angina”; and (3) “cohort” OR “cohorts” OR “follow-up” OR “followed” OR “retrospective” OR “prospective” OR “retrospectively” OR “prospectively” OR “mortality” OR “prognosis” OR “survival” OR “adverse events”. Only human studies published in English were included. The reference lists of original and review articles were manually screened as a supplementation. The final search was performed on 20 July 2019. The full-search strategy for PubMed was presented in online supplementary file 1.

Study selection
Studies were included according to the following criteria: (1) full-length article in English; (2) designed as prospective follow-up studies with a minimal follow-up duration of 1 year; (3) included at least 100 patients with ACS; (4) anxiety assessed within 3 months of the onset of ACS as exposure of interest; (5) documented the incidence of mortality (all cause or cardiovascular) and/or major adverse cardiovascular events (MACEs) in patients with and without anxiety at baseline and (6) reported the multivariable adjusted risk ratios (RRs) and their corresponding 95% CIs for mortality and/or MACEs outcomes in patients with anxiety compared with those without anxiety. MACEs were defined as a composite outcome of cardiac death, non-fatal MI, cardiac rehospitalisation, recurrence of ACS and repeated coronary revascularisation. The diagnosis of anxiety was consistent with the criteria of the original articles. For repeated reports of the same cohort, the latest studies with the longest follow-up duration were included.

Data extraction and quality evaluation
Two authors independently performed the literature search, data extraction and quality assessment. If discrepancies occurred, they were resolved by consultation with the corresponding author. A predefined form was used for data extraction. The extracted data included: (1) first author, location and design of the study; (2) number, mean age, gender and diagnosis of the patients; (3) diagnostic tools for anxiety and number of patients with anxiety; (4) follow-up durations, outcomes reported, number of patients with outcomes and variables adjusted and (5) outcome data for the mortality/ MACEs risk in patients with ACS with anxiety compared with those without anxiety as presented in RRs and 95% CIs. Study quality was evaluated using the Newcastle-Ottawa Scale.31 This scale ranges from 1 to 9 stars and assesses quality of the individual study according to the following three aspects: selection of the study groups; comparability of the groups and ascertainment of the outcome of interest.

Statistical analyses
Data of RRs and their corresponding SEs were estimated from 95% CIs or p values, and then logarithmically transformed to stabilise the variance and to normalise the distribution of the data.30 For the two studies where the OR was described,9 25 we converted data to RR for meta-analysis (RR=OR/(1−pRef)+(pRef×OR)) as previously described,32 where pRef is the prevalence of the outcome in the reference group. The Cochrane’s Q test was performed to evaluate the heterogeneity among studies, as well as the calculation of the I² statistic.33 An I²>50% indicated significant heterogeneity. A random-effect model was used for the meta-analysis since this model could incorporate the potential heterogeneity of the included studies and provide a more generalised result.30 Sensitivity analyses, which remove studies one at a time, were performed to evaluate the stability of the results.34 Subgroup analysis was performed to evaluate whether subtypes of ACS or adjusting for depression influenced the results. Risk of publication bias was assessed by funnel plots, complemented with the Egger regression asymmetry test.35 RevMan (V.5.1; Cochrane Collaboration, Oxford, UK) and STATA software (V.12.0; StataCorp) were used for the statistical analyses.
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Overall, this meta-analysis included 39,038 patients with ACS.9–25 All studies were performed in countries from North America or Europe,9–15 17–20 23 24 except two studies from Asia.16 25 Eleven studies included patients with MI,9–12 15–20 23 while the others included patients with ACS with MI and UA.13 14 21 22 24 25 The numbers of patients included in each study varied from 193 to 26,641. The mean ages of the included patients ranged from 59 to 67 years, with the percentage of male patients varying from 66% to 100%. Various symptoms-based anxiety evaluation tools were used to diagnose anxiety. The proportions of patients with anxiety ranged from 9.1% to 58.2%. Follow-up durations varied from 1 to 10 years after onset of acute coronary events. Potential factors, including demographic characteristics, comorbidities, risk factors for CAD, coronary lesion characteristics and other psychological factors such as depression, that may have confounded the association between anxiety and prognosis after ACS, were adjusted to various extents. The quality of the included studies was generally good, as evidenced by the Newcastle-Ottawa Scale, varying from 7 to 9 stars (online supplementary file 2).

### Association between anxiety and mortality in patients with ACS

The association between anxiety and mortality outcomes for patients with ACS was evaluated in 10 prospective cohorts.10 12 14–17 20 21 23 24 Moderate heterogeneity was detected (p for Cochrane’s Q test=0.22, I²=25%). Meta-analysis with a random-effect model showed that anxiety was independently associated with increased mortality risk after acute coronary events (adjusted RR 1.21, 95% CI 1.07 to 1.37, p=0.002; figure 2A). Sensitivity analysis, conducted by excluding one study at a time, retrieved similar results (RR 1.17–1.25, p all <0.05). Subgroup analyses showed that the association between anxiety at baseline and mortality risk in ACS was not significant in studies after adjusting for depression (adjusted RR 0.88, 95% CI 0.66 to 1.17, p=0.37, I²=0%), but remained significant in studies without adjusting for depression (adjusted RR 1.28, 95% CI 1.17 to 1.40, p<0.001, I²=0%; figure 2B). These findings indicate that depression may confound the association between anxiety and mortality risk after acute coronary events (p for subgroup difference=0.01). Moreover, the significant association between anxiety and increased mortality risk was observed in studies including MI patients only (adjusted RR 1.24, 95% CI 1.09 to 1.41, p<0.001, I²=29%), but not in studies with all subtypes of ACS (adjusted RR 0.96, 95% CI 0.65 to 1.40, p=0.82, I²=0%; figure 2C). However, the results between the subgroups were not significantly different (p for subgroup difference=0.20).

### Association between anxiety and MACEs in patients with ACS

Pooled results with eight prospective cohort studies9 11 13 17–19 22 25 showed that anxiety was independently associated with increased risk of MACEs after acute coronary events (adjusted RR 1.47, 95% CI 1.24 to 1.74, p<0.001; figure 3A) with moderate heterogeneity (p for Cochrane’s Q test=0.16, I²=33%). Sensitivity analyses conducted by excluding one study at a time retrieved similar results (RR 1.40–1.54, p all <0.05). Subgroup analyses showed that, although the association between anxiety and increased risk of MACEs remained significant in studies after adjusting for depression (adjusted RR 1.25, 95% CI 1.04 to 1.52, p=0.02, I²=0%), the strength of the association was attenuated compared with studies without adjusting for depression (adjusted RR 1.66, 95% CI 1.31 to 2.10, p<0.001, I²=29%; p for subgroup difference=0.07, figure 3B). Moreover, the significant association between anxiety and increased risk of MACEs after acute coronary events was significant in studies with patients with MI (adjusted RR 1.46, 95% CI 1.20 to 1.78 p<0.001, I²=17%),
Table 1  Characteristics of the included prospective follow-up studies

| Author year   | Country        | ACS type | No of patients | Men age | Male | Diagnosis of anxiety | No of patients with anxiety | Follow-up duration | Outcomes reported (no of cases) | Definition of MACE | Variables adjusted |
|---------------|----------------|----------|----------------|---------|------|----------------------|----------------------------|---------------------|-------------------------------|-------------------|--------------------|
| Frasure-Smith, 1995  | Canada | MI | 222 | 59.8 | 77.9 | STAI | 57 (25.7) | 1 | MACE (48) | Cardiac death, non-fatal MI, CV hospitalisation | Age, gender, previous MI, Killip class, and CVD medications |
| Frasure-Smith, 2003  | Canada | MI | 896 | 59.4 | 74.1 | STAI | NA | 5 | CV mortality (121) | NA | Age, gender, education, smoking, previous MI, LVEF, Killip class, secularisation therapy, and CVD medications |
| Strik, 2003  | The Netherlands | MI | 318 | 58 | 100 | SCL-90 | 185 (58.2) | 3.4 | MACE (25) | Cardiac death, non-fatal MI | Age and LVEF |
| Dickens, 2008  | UK | MI | 588 | 60 | 70.4 | HADS-A | 140 (23.8) | 6.7 | CV mortality (32) | NA | Age, gender, previous MI, Killip class, and CVD medications |
| Parker, 2010  | Australia | ACS | 489 | 67.2 | 71.3 | DSM-IV | 107 (21.9) | 1 | MACE (146) | Cardiac death, CV hospitalisation, and revascularisation | Age, LVEF, stroke history, CABG procedure at baseline, and DM |
| Doyle, 2012  | Ireland | ACS | 598 | 62.8 | 75.5 | HADS-A | NA | 8 | All-cause mortality (121) | NA | Age, sex, private health insurance, DM, ever smoker, prior ACS, HTN, reperfusion, and length of hospital stay |
| Wrenn, 2013  | USA | MI | 1968 | 60.2 | 69.4 | STPI | 179 (9.1) | 10 | All-cause mortality (525) | NA | Age, sex, BMI, marital status, race, educational attainment, smoking, previous MI, CHF, DM, HTN, noncardiac comorbidities, CV medications, social status, and alcohol consumption |
| Roest 2013  | The Netherlands | MI | 418 | 59 | 81.1 | HARS | 50 (12.0) | 3.8 | MACE (50) | Cardiac death, non-fatal MI | Age, gender, cardiac history, and LVEF |
| Hosseini, 2014  | Iran | MI | 285 | 59.1 | 69.1 | STAI | 145 (60.9) | 5 | All-cause mortality (91) | NA | Age, gender, smoking, alcohol consumption, previous MI, HTN, DM, thrombolysis therapy and depression |
| Larsen, 2014  | Denmark | MI | 896 | 67 | 69.2 | HADS-A | 211 (23.5) | 2.6 | MACE (239) and all-cause mortality (84) | Cardiac death, non-fatal MI, CV hospitalisation | Age, gender, smoking, previous MI, HTN, DM, CV medication, physical activity and depression |

Continued
| Author year | Country | ACS type | No of patients | Men age | Male | Diagnosis of anxiety | No of patients with anxiety | Follow-up duration | Outcomes reported (no of cases) | Definition of MACE | Variables adjusted |
|-------------|---------|----------|----------------|---------|------|---------------------|--------------------------|------------------|--------------------------------|-----------------|-------------------|
| Van Beek, 2016 | The Netherlands | MI | 193 | 62.1 | 65.7 | CAQ | NA | 4.3 | MACE (77) | Cardiac death, CV hospitalisation | Age, gender, LVEF, previous MI and depressive symptoms |
| Smeijers, 2017 | USA | MI | 2176 | 60.1 | 70.8 | STPI | NA | 10 | All-cause mortality (580) | NA | Age, gender, race/ethnicity, marital status, education, income, smoking status, alcohol consumption, BMI, usual physical activity, medical history of MI, HTN, DM and CV medications |
| Ossola, 2018 | Italy | ACS | 266 | 61.6 | 81.5 | HADS-A | NA | 2 | MACE (57) | Cardiac death, CV hospitalisation, recurrence of ACS, and revascularisation | Age, gender, GRACE score, and depression |
| de Jager, 2018 | The Netherlands | ACS | 528 | 63.3 | 75.9 | HADS-A | 121 (22.9) | 10 | All-cause mortality (134) | NA | Age, gender, cardiac history, indication for PCI, HTN, TC, DM, family history of CAD, multivessel disease, smoking and depression |
| Xia, 2019 | China | ACS | 647 | 63.8 | 68.4 | GAD-7 | 68 (10.5) | 1 | MACE (49) | Cardiac deaths, CV hospitalisation, non-fatal MI | Age, gender, LVEF, smoking, physical activity, DM, HTN, HC and PAD |
| Lissaker, 2019 | Sweden | MI | 26641 | 61.7 | 73.5 | EQ-5D | 5699 (21.4) | 4.3 | CV mortality (1038) | NA | Age, gender, education, smoking, previous MI, LVEF, Killip class, revascularisation therapy, CVD medications and discharged characteristics |
| Tran, 2019 | USA | ACS | 1909 | 61 | 66.7 | GAD-7 | 198 (10.4) | 2 | All-cause mortality (580) | NA | GRACE risk score, symptoms of depression, SCr, and DBP at admission |

TC, total cholesterol; HADS-A, hospital anxiety and depression scale-anxiety subscale; CHF, chronic heart failure; DSM-IV, diagnostic and statistical manual of mental disorders 4th ed. ACS, acute coronary syndrome; BMI, body mass index; CAD, coronary artery disease; CAQ, Cardiac Anxiety Questionnaire; CHF, chronic heart failure; CV, cardiovascular; CVD, cardiovascular disease; DBP, diastolic blood pressure; DM, diabetes mellitus; DSM-IV, diagnostic and statistical manual of mental disorders 4th ed; EQ-5D, European Quality of Life Five Dimensions questionnaire; GAD-7, generalised anxiety disorder scale-7; GRACE, Global Registry of Acute Coronary Events; HADS, Hospital Anxiety and Depression Scale; HADS-A, hospital anxiety and depression scale-anxiety subscale; HARS, Hamilton Anxiety Rating Scale; HC, hypercholesterolaemia; HTN, hypertension; LVEF, left ventricular ejection fraction; MACE, major adverse cardiovascular events; MI, myocardial infarction; NA, not available; NOS, the Newcastle-Ottawa Scale; PAD, peripheral artery disease; PCI, percutaneous coronary intervention; SCL-90, 90-item Symptom Check List; SCr, serum creatinine; STAI, State-Trait Anxiety Inventory; STPI, State-Trait Personality Inventory; TC, total cholesterol.
and in studies with all subtypes of patients with ACS (adjusted RR 1.49, 95% CI 1.01 to 2.18, p=0.04, I²=65%; figure 3C). The results between the subgroups were not significantly different (p for subgroup difference=0.94).

**Publication bias**

Funnel plots of the meta-analysis of the association between anxiety and mortality risk in patients with ACS: (B) subgroup analyses after adjusting for depression and (C) subgroup analyses according to the subtypes of ACS. ACS, acute coronary syndrome.

**Discussion**

Results of our meta-analysis showed that patients with ACS with anxiety at baseline have a 21% greater risk of mortality and 47% higher risk of MACEs compared with those without anxiety. However, subgroup analyses showed that concurrent depression may be a significant modifier of the association between anxiety and adverse outcomes underlining the validity of our meta-analyses for the two outcomes (p=0.527 and 0.299, respectively).
in patients with ACS. Specifically, association between anxiety and increased risk of MACEs in patients with ACS was significantly attenuated in studies that adjusted for depression, and the association between anxiety and mortality risk also became insignificant in these studies. Taken together, these results suggest that patients with ACS with anxiety at baseline have increased risk of mortality and MACEs during follow-up. However, it seems that at least part of the association may be confounded by concurrent depressive symptoms in these patients.

Two previous meta-analyses have been published that evaluated the association between anxiety and prognosis in post-MI patients. Our study has the following important clinical implications compared with the previous ones. First, we expanded the previous study population to include patients with ACS. This is significant because patients with ACS tend to have similar severity of clinical symptoms, complications and exposure to invasive procedures for diagnosis and treatment, which lead to high prevalence of anxiety. Second, we limited our inclusion criteria to prospective follow-up studies with at least 100 patients who were followed for at least 1 year. Moreover, only the adjusted RRs for the associations were included, accounting for potential confounding factors, including demographic factors, CAD risk factors, comorbidities, severity of the coronary lesions and CAD treatment.
Therefore, our study provides a more reliable result regarding the association between anxiety and prognosis in patients with ACS. Third, by including recently published studies we were able to explore the potential of depression underlying the association between anxiety and poor prognosis in patients with ACS. We found that the potential association between anxiety and increased risk of mortality and MACEs was significantly attenuated or even became insignificant in studies, which had adjusted for depression in their analyses. Previous studies have confirmed that depression is a stronger predictor of poor prognosis in post-MI patients, which may confer a 2-fold to 2.7-fold increased risk of MACEs. Moreover, since anxiety and depression are highly correlated with each other, depression may be a more important predictor of prognosis than anxiety. During the screening and management of affective disorders in patients with ACS, more efforts should be made to target depression rather than anxiety. However, considering the overlap of clinical manifestations, screening instruments, and pharmacological treatments for anxiety and depression, it would be important to manage emotional disorders in patients with acute coronary events in clinical practice.

Some implications could be summarised for clinical practice and future research in this field. The potential pathophysiological basis for the association between anxiety and poor prognosis in patients with ACS may include the following: anxiety is a stressor that might pose a further burden on cardiac function. Anxiety might also affect sleep quality, causing disruption of diurnal cycles, thus putting pressure on metabolic pathways leading to further cardiovascular damage. For those diagnosed with anxiety, prescribing anxiolytic drugs may also lead to suppressed respiration and thus further diminish oxygen in the heart. A ‘U-shaped’ association between anxiety and clinical outcomes has been noticed in some clinical settings. Patients with mild anxiety may have better compliance, which therefore leads to better clinical prognosis. However, because no data on the severity of anxiety (or dose–response data) was available in the included studies, future studies are needed to investigate whether the association between anxiety and outcomes in patients with ACS varied according to the severity of anxiety. Interestingly, results of subgroup analysis demonstrated that depression may partially confound the association between anxiety and poor prognosis in ACS. The potential mechanisms remain unknown. In view of the common coexistence of anxiety and depression in patients with affective disorders, these findings suggest that management of depression in these patients is important.

Nine different instruments were used as the evaluation tools for anxiety in the included studies. All of the included studies diagnosed anxiety as a binary variable, and the proportions of patients with anxiety at baseline varied considerably. These findings indicate that the diagnostic efficacy and cut-offs for variables anxiety evaluating tools may not be comparable, which may contribute to the heterogeneity. However, since subgroup data (or dose–effect analysis) regarding the severity of anxiety was not reported in either of the included studies, we were unable to determine the potential non-linear association between anxiety and outcomes in patients with ACS. Future studies are needed in this regard.

Despite the clinical implications of our findings, we acknowledge the following limitations of our study. First, the numbers of included studies were limited for the meta-analyses of both MACEs and mortality outcomes, and were even further limited for subgroup analyses. Therefore, results of the subgroup analyses should be interpreted cautiously and validated in the future. Second, although no statistical heterogeneity was detected for our meta-analysis, the instrument and timing of anxiety evaluation may have affected the results, which should be evaluated in the future. Third, we included studies that reported both cardiac mortality and all-cause mortality, and definitions of MACEs varied across the studies, which may have also introduced biases. Fourth, our study is a meta-analysis of observational studies, which do not provide a causative association between anxiety and poor
outcomes in patients with ACS. Clinical studies evaluating the potential benefits of treatments targeting anxiety in patients with ACS are warranted. Finally, we only included publications in English, and relevant studies published in other languages were omitted.

In conclusion, results of this meta-analysis showed that patients with ACS with anxiety at baseline have an increased risk of mortality and MACEs during a follow-up. At least part of this association may be confounded by concurrent depressive symptoms in these patients. More studies are needed to evaluate the potential interactions between anxiety and depression on clinical outcomes in patients with ACS and to determine the optimal management of these potential affective disorders in patients with acute coronary events.

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Data availability statement Data are available on reasonable request. The datasets generated and analysed during the present study are available from the corresponding author on reasonable request.

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