Data-driven stability analysis of switched linear systems with Sum of Squares guarantees

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Abstract: We present a new data-driven method to provide probabilistic stability guarantees for black-box switched linear systems. By sampling a finite number of observations of trajectories, we construct approximate Lyapunov functions and deduce the stability of the underlying system with a user-defined confidence. The number of observations required to attain this confidence level on the guarantee is explicitly characterized.

Our contribution is twofold: first, we propose a novel approach for common quadratic Lyapunov functions, relying on sensitivity analysis of a quasi-convex optimization program. By doing so, we improve a recently proposed bound. Then, we show that our new approach allows for extension of the method to Sum of Squares Lyapunov functions, providing further improvement for the technique. We demonstrate these improvements on a numerical example.

Keywords: Stability analysis of switched linear systems, Quasi-convex optimization, Data-driven analysis, Scenario approach, Sum of Squares techniques.

1. INTRODUCTION

Hybrid systems, whose dynamics include continuous and discrete parts, allow to model a wide range of complex systems, including cyber-physical systems which have been a subject of growing interest recently (Lee (2015), Tabuada et al. (2014)). An important family of hybrid systems are switched linear systems, consisting of several linear continuous state models, the modes, together with a discrete rule deciding the switching among them. Formally, let $M = \{1, 2, \ldots, m\}$, and $\mathcal{M} = \{A_i, \ i \in M\}$ be a finite set of $m$ matrices in $\mathbb{R}^{n \times n}$, the modes. We consider systems of the form

$$x_{k+1} = A_\tau(k) x_k$$

where $x_k \in \mathbb{R}^n$ is the state, $k \in \mathbb{N}$ is the time index, and $\tau \in M^\mathbb{N}$ is the switching sequence taking values in $M$. Switched systems can for instance be used to model linear systems subject to time-dependent uncertainty or multi-controller switching schemes (Liberzon (2003), Shorten et al. (2007)).

A system of the form (1) is uniformly asymptotically stable if it possesses the following property:

$$\forall \tau \in M^\mathbb{N}, \forall x_0 \in \mathbb{R}^n, \|x_k\| \rightarrow 0.$$  

This property is characterized by the joint spectral radius (JSR) of $\mathcal{M}$ (Jungers (2009)):

$$\rho(\mathcal{M}) = \lim_{k \rightarrow \infty} \max_{i_1, \ldots, i_k} \|A_{i_1} \cdots A_{i_k}\|^{\frac{1}{k}} : A_{i_j} \in \mathcal{M}.$$  

Indeed, given a finite set of matrices $\mathcal{M}$, the corresponding switched dynamical system is uniformly asymptotically stable if and only if $\rho(\mathcal{M}) < 1$ (Jungers (2009), Cor. 1.1). Because switched linear systems do not inherit the properties of their modes, their stability analysis is difficult: deciding whether their JSR is smaller than 1 is NP-hard (Blondel et al. (1999)), and many tools have therefore been designed in the past decades for their analysis. Most of these tools are model-based, build on notions from Lyapunov and invariant set theory, and provide deterministic guarantees when the number of modes is finite (Lin et al. (2009), Blanchini et al. (2007), Liberzon et al. (2004)). However, in many practical applications, a closed-form model may not be available (think for example of self-driving cars, complex robotic systems, or smart grids applications). Hence, stability analysis based on information obtained via a sufficient number of simulations (or observations of the unknown system) has received increasing attention lately (Kozarev et al. (2016), Blanchini et al. (2017), Kenanian et al. (2019), Smarra et al. (2020), Wang et al. (2021)). A first approach consists in using model-based stability analysis techniques after recovering the model from the simulations. However, since the identification problem for switched systems is NP-hard (Lauer (2016)), we bypass it in our approach. We refer to our framework as data-driven analysis of black-box dynamical systems. It has been increasingly understood that one can obtain effective control methods, such as Lyapunov functions or invariant sets, despite the complexity of the underlying system and the lack of a closed-form description of the model.

In Kenanian et al. (2019), a probabilistic stability guarantee for black-box switched linear systems is pro-
vided, based on Common Quadratic Lyapunov Functions (CQLFs) decreasing on a finite number \( N \) of observations of their trajectories. The method proceeds by deriving a guarantee on the measure of the subset of the state space on which the CQLF is not decreasing using scenario optimization theory (Campi et al. (2018), Calafiore (2010)), and then transforming this guarantee into a global stability certificate with a certain confidence level, using geometric properties of switched systems.

Relying on CQLFs is conservative for switched linear systems, since some stable switched systems do not admit a CQLF (Lin et al. (2009)). For this reason, in this paper, we derive a broader class of probabilistic stability guarantees for black-box switched linear systems. We generalize the data-driven analysis technique to a more complex class of Lyapunov functions decreasing along the observed trajectories, namely Sum of Squares (SOS) Lyapunov functions (Parrilo et al. (2008), Jungers (2009), Section 2.3.7).

Our main contribution consists in the derivation of a probabilistic upper bound on the JSR of a black-box system, with a user-defined confidence level, via a SOS Lyapunov function of arbitrary degree \( d \). The trade-off between the number of observations \( N \) and the tightness of the upper bound is explicitly quantified. Because the geometric tools used in Kenanian et al. (2019) do not generalize in a straightforward way to SOS Lyapunov functions of higher degrees, we develop another approach, based on sensitivity analysis of a quasi-convex optimization problem (Bertsekas (2016)). Hence, our new approach does not rely on classical chance-constrained theorems usually underlying the scenario approach. Additionally, we thereby also obtain a novel upper bound in the quadratic case, that can be compared with the one obtained in Kenanian et al. (2019). In the white-box setting, approximations of the JSR using SOS Lyapunov functions are tighter (Parrilo et al. (2008)), and we will observe that this generalizes to the black-box setting.

The outline of this paper is as follows. This section ends with the notation, and Section 2 contains preliminary results about the approximation of the JSR using CQLFs as well as the introduction of the problem studied. Section 3 presents our new approach relying on sensitivity analysis, and our novel probabilistic upper bound on the JSR in the quadratic case. In Section 4, we leverage this approach in order to obtain our bound relying on SOS functions. Finally, we illustrate the performance of our bounds in Section 5 with a numerical experiment.

The notation used in this paper is as follows. For a square matrix \( P \), \( P > 0 \) (resp. \( \geq 0 \)) states that \( P \) is positive definite (resp. semi-definite). We denote the set of real symmetric matrices of size \( n \) by \( \mathcal{S}^n \), and the set of positive definite matrices by \( \mathcal{S}^n_+ \). For \( P \in \mathcal{S}^n_+ \), we define the ellipsoidal vector norm associated to \( P \) as: \( \| x \|_P := \sqrt{x^T P x}. \) ||x|| := \sqrt{x^T x} denotes the classical \( \ell_2 \)-norm of \( x \in \mathbb{R}^n \).

Our notation concerning measure-theoretic concepts is taken from Ledoux et al. (1991). \( \mathcal{S} = \{ x \in \mathbb{R}^n : ||x|| = 1 \} \) denotes the unit sphere centered at the origin in \( \mathbb{R}^n \). \( \sigma^{n-1} \) denotes the uniform spherical measure, satisfying \( \sigma^{n-1}( \mathcal{S} ) = 1 \). \( \mu_S \) denotes the uniform measure on \( M = \{ 1, 2, ..., m \} \) for \( m \in \mathbb{N}_0 \) and \( \mu_M = \otimes^m \mu_M \) denotes the uniform product measure on \( M^t = \otimes^t M \), for \( t \in \mathbb{N}_0 \).

Finally, \( \mu_t = \sigma^{n-1} \otimes \mu_M^t \) denotes the uniform measure on \( Z_t = \mathcal{S} \times M^t \).

### 2. PRELIMINARIES

#### 2.1 Approximation of the JSR using quadratic forms

A framework to approximate the JSR of switched linear systems and hence to analyze their stability is provided by CQLFs:

**Proposition 1.** (Jungers (2009), Prop. 2.8) Consider a finite set of matrices \( M \in \mathbb{R}^{n \times n} \). If there exists \( \gamma \geq 0 \) and \( P \in \mathcal{S}^n_+ \) such that

\[
\forall A \in M, A^T PA \leq \gamma^2 P
\]

then \( \rho(M) \leq \gamma \).

When \( \gamma < 1 \), \( \| \cdot \|_P \) is a CQLF and can serve as stability certificate for System (1). In the spirit of Kenanian et al. (2019) and in order to benefit from the observation of long trajectories, we consider CQLFs for the finite set \( M^t := \{ \prod_{j=1}^t A_i : A_i \in M \} \) and use the relation \( \rho(M^t) = \rho(M)^t \) (Jungers (2009), Prop. 2.5). Since the quality of the upper bound in Proposition 1 increases when \( \gamma \) decreases, we seek for the minimal \( \gamma \) such that (2) is satisfied, by considering the following optimization problem:

\[
\min_{\gamma \geq 2} \gamma
\]

s.t. \( (Ax)^T P A x \leq \gamma^2 x^T P x, \forall A \in M^t, \forall x \in \mathcal{S}^t \)

\[ P \geq I, \quad \| P \| \leq C \] for a large \( C \in \mathbb{R}_{\geq 0} \).

We denote by \((\gamma^o, P^o)\) the solution to Problem (3). While the optimal cost \( \gamma^o \) is unique, there can be several optimal \( P \). When this occurs, we apply a tie-breaking rule: \( P^o \) denotes the optimal \( P \) with the smallest condition number, since the quality of our upper bound will depend on this number. Problem (3) differs from Program (2) in three ways. First, we impose a bound on the norm of \( P \) to ensure the compactness of the set of constraints. Second, \( P \geq I \) replaces the constraint \( P > 0 \), and third, we restrict \( x \in \mathcal{S} \). These latter restrictions do not incur any conservativeness because of the homogeneity of System (1): for \( \lambda > 0 \), \( \lambda A(x) x_k \) is \( \lambda A(x) x_k \), implying for instance that the decrease of a CQLF on an arbitrary set enclosing the origin is sufficient to serve as a stability certificate.

The bound on the JSR obtained by Program (3) can, in the white-box setting, be improved by considering SOS polynomials \( p(x) \), that is polynomials admitting a decomposition \( p(x) = \sum p_i(x)^2 \) (Parrilo et al. (2008)).

Before we investigate on how this generalizes to the black-box setting, we formally present the problem addressed in this paper as well as the new data-driven approach we propose in the quadratic case.

#### 2.2 Problem statement

Problem (3) is defined by an infinite number of constraints, while we only sample a finite number \( N \) of trajectories of length \( l \) \( (x_{i,0}, x_{i,1}, ..., x_{i,l}) \) of System (1), corresponding to a finite subset of the constraints of Problem (3). These trajectories as well as an upper bound \( m \) on the number of modes of the actual system are the only available information about the system.

The trajectories are assumed to be generated from \( N \) initial states \( x_{i,0} \) drawn randomly, uniformly and independently from \( \mathcal{S} \) according to \( \sigma^{n-1} \). Then, \( N \) sequences
of l modes are drawn randomly, uniformly and independently from \( M^l \) according to \( \mu_M \) and applied to each initial condition. The set of N observations of System (1) \( \{(x_{i,0}, x_{i,1}, \ldots, x_{i,l}) \mid i = 1, 2, \ldots, N\} \) is therefore associated to a uniform sample of \( N \) \((l + 1)\)-tuples in \( Z_l^* \):

\[
\omega_N := \{(x_{i,0}, x_{i,1}, \ldots, x_{i,l}) \mid i = 1, 2, \ldots, N\} \subset Z_l.
\]

We denote \( A_{l,j} := A_{j-1}A_{j-2}\cdots A_{j-1} \) and \( j := j_1, \ldots, j_l \). For a given data set \( \omega_N \), we can define a sampled optimization problem associated to (3), that is a similar problem composed of the \( N \) constraints associated to the \( N \) observed trajectories. We denote it by \( \text{Opt}(\omega_N) \):

\[
\min_{P \geq I, \|P\| \leq C} \gamma
\]

s.t. \( \langle A_{l,j}^T P A_{l,j} \rangle x \leq \gamma \|x\|^2 P x, \forall (x, j) \in \omega_N \) (5)

Its solution is denoted by \( (\gamma(\omega_N), P(\omega_N)) \), where the tie-breaking rule is applied for \( P(\omega_N) \) if needed. Observe that Program (5) only requires the knowledge of \( A_{l,j} \), available through the observations: it does not require the knowledge of \( J \), which is not observed in our setting. Once the solution to Problem (5) is obtained, we aim to infer the solution to Problem (3) within a certain confidence.

3. A NEW DATA-DRIVEN APPROACH FOR THE JSR COMPUTATION

In this section, we provide a sensitivity analysis of the quasi-convex optimization problem (3), allowing to approximate its solution. From this, we obtain a novel probabilistic upper bound on the JSR of \( M \). An asset of this approach is the possibility to extend it to the SOS framework and hence to take advantage of the tighter approximation of the JSR allowed by SOS Lyapunov functions.

3.1 Sensitivity analysis

First, we derive a key result about the existence and the cardinality of a finite set of points such that the solutions to Problems (3) and (5) defined on this set coincide \(^1\).

**Lemma 1.** Let \( \gamma^o \) be the optimal cost of Problem (3). Then, there exists a set \( \omega_s \subset Z_l \) with \( |\omega_s| = \frac{n(n+1)}{2} + 1 \) such that \( \gamma^o(\omega_s) = \gamma^o \), where \( \gamma^o(\omega_s) \) is the optimal solution to the problem (5).

**Proof.** From a standard continuity argument (see Appendix A), it is equivalent to prove that for any \( \epsilon > 0 \), there exists a set \( \omega_s \subset Z_l \) with \( |\omega_s| = \frac{n(n+1)}{2} + 1 \) such that \( \gamma^o(\omega_s) > \gamma^o - \epsilon \). Let \( d_1 = \frac{n(n+1)}{2} \), be the number of variables of \( P \in \mathbb{S}^N \). Problem (3) is equivalent to:

\[
\min_{P \geq I, \|P\| \leq C} \gamma
\]

s.t. \( \langle A_{l,j}^T P A_{l,j} \rangle x \leq \gamma \|x\|^2 P x, \forall (x, j) \in \omega_N \) \( \in \mathbb{X}(\epsilon, \gamma) \) \( \in \mathbb{Z} \).

\[
\mathbb{X}(\epsilon, \gamma) := \{(x, j) \mid \|x\| \leq d_1 + 1, \forall (x, j) \}
\]

for any \( \omega \subset Z_l \) with \( |\omega| = d_1 + 1 \), \( \cap_{(x,j) \in \omega} \mathbb{X}(\epsilon, \gamma) \) is nonempty. Hence, by Helly’s theorem (Danzer et al. (1963)), \( \cap_{(x,j) \in \omega} \mathbb{X}(\epsilon, \gamma) \) is nonempty. As a result, \( \gamma^o \leq \gamma^o(\omega_s) = \gamma^o \), which leads to a contradiction.

Based on Lemma 1, we will show how to construct, from the set of trajectories \( \omega_N \), another set \( \omega_s \) such that \( \gamma^o(\omega_s) = \gamma^o \) with a certain probability depending on \( N \) and on how close the points of both sets are. Firstly, we introduce the notion of a spherical cap:

**Definition 1.** (Li (2011))

The spherical cap on \( S \) of direction \( c \) and measure \( \epsilon \) is defined as \( C(c, \epsilon) := \{x \in S : \|c\| \leq \epsilon\} \), where \( \epsilon(\delta) \) is given by:

\[
\epsilon(\delta) = \sqrt{1 - I^{-1}(2; n - \frac{n - 1}{2}, \frac{1}{2})}
\]

where \( I^{-1}(y; a, b) \) is the inverse regularized incomplete beta function (Majumder et al. (1973)). Its output is \( x > 0 \) such that \( I(x; a, b) = y \), where \( I \) is given by:

\[
I : \begin{cases}
R_{>0} \times R_{>0} \times R_{>0} & \rightarrow R_{>0} \\
(x, a, b) & \rightarrow I(x; a, b) = \int_0^x \frac{1}{b} \left(1 - \frac{t}{b}\right)^{a-1} dt
\end{cases}
\]

Proposition 2. Let \( \omega_N := \{(x_i, j_i) \mid i = 1, \ldots, N\} \) be a uniform random sample drawn from \( Z_l \), with \( N \geq \frac{n(n+1)}{2} + 1 \), and \( \gamma^o \) be the optimal cost of Problem (3).

For all \( \epsilon \in [0,1] \), with probability at least \( \beta(\epsilon, m, N) := 1 - \left(\frac{n(n+1)}{2} + 1\right)(1 - \frac{\epsilon}{mN}) \), there exists an \( \omega_N \) with \( \omega_N \subset \mathbb{S}(x_i, j_i), i = 1, \ldots, N \) such that:

- \( \gamma^o(\omega_N^\epsilon) = \gamma^o \), where \( \gamma^o(\omega_N^\epsilon) \) is the optimal solution to the problem (5).
- \( \text{For } i = 1, \ldots, N \text{, } \|x_i - x_i'\| \leq \epsilon(\delta) \), where \( \delta(\cdot) \) is given by (7).

**Proof.** By Lemma 1, there exists a set \( \omega_s \), with \( |\omega_s| = \frac{n(n+1)}{2} + 1 \), such that \( \gamma^o(\omega_s) = \gamma^o \). Given any \( (x_i', j_i') \in \omega_s \), let \( x_i \), \( j_i \) be drawn randomly and uniformly from \( Z_l \). Then, the probability that \( x \in C(x', \epsilon) \) and \( j = j' \) is of \( \frac{1}{mn} \). Hence, \( \mu_1(x_i, j_i) = C(x', \epsilon) \) or \( j \neq j' \) is \( \leq \frac{1 - \epsilon}{m} \) for all \( (x_i, j_i) \in \omega_N, N \geq \frac{n(n+1)}{2} \). Since these sets, for each \((x_i', j_i') \in \omega_s\), can be disjoint; the measure of their union, that is the probability that there exists no \((x_i, j_i) \in \omega_N \) such that \( x \in C(x', \epsilon) \) and \( j = j' \) for a least one \((x_i', j_i') \in \omega_s\), is smaller or equal to \( \left(\frac{n(n+1)}{2} + 1\right)(1 - \frac{\epsilon}{mn}) \). Therefore, the probability of \( \mathbb{V}(x', j') \in \omega_N \) such that to exist point \((x_i, j_i) \in \omega_N \) such that \( x \in C(x', \epsilon) \) and \( j = j' \) is at least of \( \beta(\epsilon, m, N) = 1 - \left(\frac{n(n+1)}{2} + 1\right)(1 - \frac{\epsilon}{mn}) \).

If \( x \in C(x', \epsilon) \), then \( x > \|x\| \delta(\epsilon) \) is given by (7), and \( \|x - x'\| \leq \epsilon(\delta) \). Therefore, we define \( \omega_N^\epsilon \) as follows: for the \( (x_i, j_i) \in \omega_N \) such that \( x \in C(x', \epsilon) \) and \( j = j' \) for some \((x_i', j_i') \in \omega_s, (x_i', j_i') \in \omega_s\). For all other \((x_i, j_i) \in \omega_N, (x_i, j_i) \in \omega_N^\epsilon \). Hence, with probability larger than \( \beta(\epsilon, m, N) \), \( \omega_N^\epsilon \) contains \( \omega_N \), which implies \( \gamma^o(\omega_N^\epsilon) = \gamma^o(\omega_N) = \gamma^o \). Moreover, by construction, for \( i = 1, \ldots, N, x_i \in \mathbb{S} \) and \( \|x_i - x_i'\| \leq \epsilon(\delta) \).

Given \( \beta, m, N \), we define \( \epsilon(\beta, m, N) = m^l \left(1 - \frac{2(1-\delta)}{n(n+1)+2}\right) \).
Using Proposition 2, we provide a probabilistic upper bound on $\gamma^\ast$, based on $(\gamma^\ast(\omega_N), P^\ast(\omega_N))$, the only solution we have access to:

**Proposition 3.** Let $\omega_N$ be a uniform random sample drawn from $Z_l$, with $N \geq \frac{n(n+1)}{2} + 1$; $(\gamma^\ast(\omega_N), P^\ast(\omega_N))$ be the optimal solution to Problem (5), and $(\gamma^\ast, P^\ast)$ be the optimal solution to Problem (3). For any $\beta \in (0, 1]$, let $\epsilon = m^l \left(1 - \frac{\sqrt{2(1-\beta)}}{n(n+1)+2}\right)$. Then, with probability at least $\beta$, the following holds:

$$\gamma^\ast \leq \sqrt{\gamma^\ast(\omega_N) + (\gamma^\ast(\omega_N) + B(M^l))\Delta(\epsilon)\kappa(P^\ast(\omega_N))}$$

(8)

where $\Delta(\epsilon) = \sqrt{2 - 2\delta(\epsilon)}$, $\delta(\cdot)$ is given by (7), $B(M^l) = \max_{\mathbf{A} \in M^l}||\mathbf{A}||$ and $\kappa(P) = \sqrt{\frac{\lambda_{\max}(P)}{\lambda_{\min}(P)}}$.

**Proof.** Let $\gamma = \gamma^\ast(\omega_N)$ and $P = P^\ast(\omega_N)$. By definition, $||\mathbf{A}_i||_p \leq \gamma||\mathbf{x}_i||_p \\forall (x_i, j) \in \omega_N$. Consider the Cholesky decomposition $P = L^T L$. Then (Golub 1989):

$$||x||_p = \sqrt{x^T L^T L x} = ||Lx|| \text{ and } ||L^{-1} x||_p = ||x||_p.$$  

Since $||A||_p = ||LAL^{-1}||$, $||L^{-1}||_p = ||L^{-1}||_p$. Hence, $||L||_p \leq \frac{1}{l}$, $||L^{-1}||_p \leq \frac{1}{l}$. Therefore, with probability at least $\beta$ and $\beta_1 - 1$, we have:

$$\rho(M) \leq \sqrt{\gamma^\ast(\omega_N) + (\gamma^\ast(\omega_N) + A(\epsilon))\Delta(\epsilon)\kappa(P^\ast(\omega_N))}$$

(15)

where $\Delta(\epsilon) = \sqrt{2 - 2\delta(\epsilon)}$, $\delta(\cdot)$ is given by (7) and $A(\epsilon) = \frac{\epsilon\sqrt{l}}{\sqrt{1 - \epsilon l}}$.

A proof of Proposition 4 is provided in the Appendix. We now propose our novel bound, by combining Propositions 3 and 4.

**Theorem 1.** Consider an $n$-dimensional switched linear system as in (1) and a uniform random sampling $\omega_N \subset Z_l$, where $N \geq \frac{n(n+1)}{2} + 1$. Let $(\gamma^\ast(\omega_N), P^\ast(\omega_N))$ be the optimal solution to (5), and $(\lambda^\ast(\omega_N), B^\ast(\omega_N))$ be the optimal solution to (14). Then, for any $\beta \in [0, 1)$, with probability at least $\beta$, we have:

$$\max_{\mathbf{A} \in M^l}||\mathbf{A}|| \leq \frac{\lambda^\ast(\omega_N)}{\sqrt{2\epsilon(\epsilon(\beta_1, m, N))}}$$

(16)

where $\epsilon(\cdot)$ is given by (7) and $\epsilon(\beta_1, m, N) = \frac{m^l}{1 - \sqrt{1 - \beta_1}}$.

In the bound of Proposition 3, the expression $B(M^l) = \max_{\mathbf{A} \in M^l}||\mathbf{A}||$ appears. It can be computed as (Jungers 2009, Prop. 2.7):

$$B(M^l) = \min_{\lambda \geq 0} \lambda, \text{ s.t. } ||\mathbf{A}|| \leq \lambda, \forall x \in \mathbb{S}, \forall \mathbf{A} \in M^l.$$  

(13)

Since we have no direct knowledge of $\mathbf{A} \in M^l$, we seek to obtain a probabilistic bound on $B(M^l)$ by solving the sampled problem

$$\min_{\lambda \geq 0} \lambda, \text{ s.t. } ||\mathbf{A} x|| \leq \lambda, \forall (x, j) \in \omega_N \quad (14)$$

which only requires the knowledge of the observations we have access to. Let $\lambda^\ast(\omega_N)$ denote the solution to Problem (14). By combining classical results from chance-constrained optimization and the approach proposed in Kenanian et al. (2019), we obtain the following bound:

**Proposition 4.** Consider a set of matrices $\mathcal{M}$, and a uniform random sampling $\omega_N \subset Z_l$, where $N > 2$. Let $(\gamma^\ast(\omega_N), P^\ast(\omega_N))$ be the optimal solution to Problem (14). Then, for any $\beta_1 \in [0, 1)$, with probability at least $\beta_1$, we have:

$$\max_{\mathbf{A} \in M^l}||\mathbf{A}|| \leq \frac{\lambda^\ast(\omega_N)}{\sqrt{2\epsilon(\epsilon(\beta_1, m, N))}}$$

(15)

where $\epsilon(\cdot)$ is given by (7) and $\epsilon(\beta_1, m, N) = \frac{m^l}{1 - \sqrt{1 - \beta_1}}$.

A proof of Proposition 4 is provided in the Appendix. We now propose our novel bound, by combining Propositions 3 and 4.
4. PROBABILISTIC SOS-BASED GUARANTEES

We will now extend our approach to the SOS framework. First, we recall properties of SOS Lyapunov functions in the white-box setting, and present generalized versions of Problems (3) and (5). Then, we derive a probabilistic upper bound on the JSR using sensitivity analysis. Before proceeding further, we recall the notion of symmetric algebra of a vector space:

**Definition 2.** (Jungers (2009) and Parrilo et al. (2008)) Let $x \in \mathbb{R}^n$. Let $D$ denote the number of different monomials of degree $d$:

$$D = \binom{n + d - 1}{d}.$$ 

(17)

The $d$-lift of $x$, denoted by $x^{[d]}$, is the vector in $\mathbb{R}^D$, indexed by all the possible exponents $\alpha$ of degree $d$:

$$x^{[d]}_\alpha = \sqrt{\alpha!} x^\alpha,$n

(18)

where $\alpha = (\alpha_1, ..., \alpha_n)$ with $|\alpha| = d$, and $\alpha! = \binom{d}{\alpha_1, ..., \alpha_n}$ is the multinomial coefficient. The $d$-lift of the matrix $A$ is the matrix $A^{[d]} \in \mathbb{R}^{D \times D}$ associated to the linear map $A^{[d]} : x^{[d]} \mapsto (Ax)^{[d]}$.

With the lift as defined in (18), one has:

$$||x^{[d]}|| = ||x||^d.$$ 

(19)

We are now ready to introduce SOS forms to approximate the JSR of $\mathcal{M}$.

**Proposition 5.** (Parrilo et al. (2008), Thm. 2.3) A homogeneous multivariate polynomial $p(x)$ of degree $2d$ is a SOS polynomial if and only if for some $P \in S_+^{n \times n}$, where $D$ is given by (17),

$$p(x) = (x^{[d]})^T P x^{[d]}.$$ 

We can obtain tighter approximations of the JSR using SOS functions due to the following property:

**Proposition 6.** (Jungers (2009), Thm. 2.13) Consider a finite set of matrices $\mathcal{M} \in \mathbb{R}^{n \times n}$. If there exists $\gamma \geq 0$ and $P \in S_+^{n \times n}$ such that

$$\forall A \in \mathcal{M}^I, \forall x \in \mathbb{S}, (Ax)^{[d]}^T P (Ax)^{[d]} \leq \gamma^{2d} (x^{[d]})^T P x^{[d]}$$

then $\rho(\mathcal{M}) \leq \gamma$.

The restriction of $x$ to $\mathbb{S}$ is again due to the homogeneity of System (1). SOS functions hence become SOS Lyapunov functions when $\gamma < 1$. We generalize Problem (3) as follows:

$$\begin{align*}
\min_{\gamma \geq 0} & \gamma \\
\text{s.t.} & \ (Ax)^{[d]}^T P (Ax)^{[d]} \leq \gamma^{2d} (x^{[d]})^T P x^{[d]}, \forall A \in \mathcal{M}^I, \forall x \in \mathbb{S} \\
& \ P \succeq I, \quad ||P|| \leq C \text{ for a large } C \in \mathbb{R}_+.
\end{align*}$$

(20)

Its solution is denoted by $(\gamma_{\text{sos}}^{\rho(\mathcal{M})}, P_{\text{sos}}^{\rho(\mathcal{M})})$, with the same tie-breaking rule applied to $P_{\text{sos}}^{\rho(\mathcal{M})}$, and the latter is not unique. When $d = 1$, we recover Problem (3). We consider the following associated sampled problem, denoted by Opt$_{\text{sos}}(\omega_N)$ and with solution $(\gamma_{\text{sos}}^*(\omega_N), P_{\text{sos}}^*(\omega_N))$:

$$\begin{align*}
\min_{\gamma \geq 0} & \gamma \\
\text{s.t.} & \ (A_jx)^{[d]}^T P (A_jx)^{[d]} \leq \gamma^{2d} (x^{[d]})^T P x^{[d]}, \forall (x, j) \in \omega_N \\
& \ P \succeq I, \quad ||P|| \leq C \text{ for a large } C \in \mathbb{R}_+.
\end{align*}$$

(21)

Applying Proposition 2 to Problem (20) (only the dimension of the optimization program changes from $\frac{n(n+1)}{2}$ to $\frac{D(D+1)}{2}$), we obtain a similar covering of the sampling space, with a confidence level $\beta_{\text{sos}, m, N} = 1 - \left(\frac{D(D+1)}{2}\right) + 1(1 - \frac{\epsilon}{\sqrt{N}})^N$. Combining this with Proposition 4, we obtain a SOS-based probabilistic upper bound on $\rho(\mathcal{M})$:

**Theorem 2.** Consider an $n$-dimensional switched linear system as in (1). Let $D = \binom{n + d - 1}{d}$ be given by (17), and $\omega_N \subset \mathbb{Z}_1$ be a uniform random sampling, with $N \geq \frac{D(D+1)}{2} + 1$. Let $(\gamma^*(\omega_N), P^*(\omega_N))$ be the optimal solution to (21) and $\lambda^*(\omega_N)$ be the optimal solution to (14). For any $\beta \in [0, 1)$ and $\beta_1 \in [0, 1)$, let $\epsilon_{\text{sos}} = \mu\left(1 - \frac{1}{2\sqrt{D(D+1)+1}}\right)$ and $\epsilon_1 = \frac{\mu}{2}(1 - \sqrt{1 - \beta})$. Then, with probability at least $\beta + \beta_1 - 1$, we have:

$$\rho(\mathcal{M}) \leq \frac{\sqrt{n\gamma(\omega_N)^{dl} + (\gamma(\omega_N)^{dl} + A(\epsilon_1)^d)f(d, \epsilon_{\text{sos}})}\kappa(P^*(\omega_N))}{\kappa(A(\epsilon_1))}$$

where $A(\epsilon_1) = \frac{\lambda^*(\omega_N)}{\lambda^*(\omega_N)}$,

$$f(d, \epsilon_{\text{sos}}) = \sqrt{D}((1 + \Delta(\epsilon_{\text{sos}}))^d - 1 - (1 - \frac{1}{\sqrt{N}})\Delta(\epsilon_{\text{sos}})^d),$$

$$\Delta(\epsilon_{\text{sos}}) = \sqrt{2} - 2d(\epsilon_{\text{sos}})$$

and $\cdot$ is given by (7).

When $d = 1$, this theorem reduces to Theorem 1 as expected. A proof of Theorem 2 is provided in the Appendix. Since $\epsilon_1$ does not depend on $d$, the minimum number of samples required to obtain a finite bound is the same as in the quadratic case.

5. EXPERIMENTAL RESULTS

In this section, we conduct an experiment in order to illustrate the improvement of our quadratic bound compared to the one of Kenanian et al. (2019), and the performance of our SOS bound compared to the quadratic bounds. We consider a stable system for which no CQFL exists (Parrilo et al. (2008), Ex. 2.8): $\mathcal{M} = \{\begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}, \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix}\}$. The true JSR of this set is equal to $1$, but the best quadratic bound (as given by (3)) is equal to $\sqrt{2}$, while one can obtain better bounds using SOS functions. First, using $N$ observations, for $N$ ranging between $10$ and $10000$, we solve Problems (14), (5) in the quadratic case, and (21) in the SOS case, for a SOS function of degree $2d = 4$. Once $\gamma(\omega_N)$ is obtained, we apply the tie-breaking rule in order to obtain a feasible $P$ of minimal condition number. Then, we compute the bounds of Theorems 1 and 2 as well as the bound of Kenanian et al. (2019). $\beta$ and $\beta_1$ are fixed to $0.95$, for a global confidence level of $\beta + \beta_1 - 1 = 0.9$, and the length of the trace is set to $1$. The average of the bounds is computed over $10$ runs for each value of $N$.

Fig. 1 presents the evolution of the different upper bounds with the number of samples. We observe that for a large enough number of samples, the SOS bound outperforms the quadratic bound in the white-box setting just like it does in the black-box setting. We also observe that while the bound of Kenanian et al. (2019) is tighter than our quadratic bound for a sufficiently high number of samples, our bound is better for a small number of samples. The fact that our bound is better for a small number of samples...
makes it even more advantageous when longer traces are used. Indeed, for larger traces, because the probability space to be sampled is larger, the performance of the bounds decreases, but faster for the bound of Kenanian et al. (2019) than for our bound.

6. CONCLUSION

We have introduced a new data-driven approach providing probabilistic upper bounds on the JSR of an unknown system, based on a finite number of observations of trajectories. Our approach relies on sensitivity analysis of a quasi-convex optimization problem, related to the construction of approximate Lyapunov functions. Based on this sensitivity analysis, we firstly derived a new bound using common quadratic Lyapunov functions, and then generalized our method to SOS Lyapunov functions. This approach is well known to improve the quality of approximation of the JSR in the white-box setting. We demonstrated on a numerical experiment that it indeed improves the JSR approximation in the black-box setting too.

For future work, we plan to investigate on the effect of adaptive sampling on the performance of our bounds: we believe that our analysis can be improved by sampling points a posteriori; and to compare our results with the ones that can be obtained by first identifying the model. We also leave for further work the generalisation of data-driven stability analysis of switched linear systems to other Lyapunov functions, for instance to the path-complete Lyapunov framework.

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Appendix A. CONTINUITY ARGUMENT

We prove the equivalence between the two following statements:

*(B1)* There exists a set \([\omega] \subset Z_t\) with \(|\omega| = d_l + 1\) such that \(\gamma^*(\omega) = \gamma^*\).

*(B2)* For any \(\epsilon > 0\), there exists a set \([\omega] \subset Z_t\) with \(|\omega| = d_l + 1\) such that \(\gamma^*(\omega) > \gamma^* - \epsilon\).

Clearly, if \((B1)\) holds then \((B2)\) holds. Suppose that \((B2)\) holds. Then, we can construct a sequence \(\{(\omega_k, \epsilon_k)\}\) where \(\omega_k \subset Z_t\), \(|\omega| = \frac{2(n+1)}{2} + 1\) and \(\epsilon_k + 1 - \epsilon_k\) such that \(\epsilon_k \to 0\) and \(\gamma^*(\omega_k) > \gamma^* - \epsilon_k\). Taking the limit on both sides yields \(\lim_{k \to \infty} \gamma^*(\omega_k) \geq \gamma^\circ\), implying \(\lim_{k \to \infty} \gamma^*(\omega_k) = \gamma^\circ\) since \(\forall \omega_k \in Z_t\), \(\gamma^*(\omega_k) \leq \gamma^\circ\).

By compactness of the set \(Z_t\), \(\{\omega_k\}\) (or a subsequence of \(\{\omega_k\}\)) converges to a limit \(\omega = \lim_{k \to \infty} \omega_k\). We now show that that \(\gamma^*(\omega)\) is continuous, which implies that \(\gamma^*(\omega) = \lim_{k \to \infty} \gamma^*(\omega_k) = \gamma^\circ\) so that \((S1)\) holds. We prove that for a given \(\epsilon > 0\) defining \(\delta := \frac{2n(\epsilon P)}{\max_{A \in \mathcal{A}^d} ||A||}\), if \(\Delta \omega \in (\delta B) \times \ldots \times (\delta B)\), then \(\gamma^*(\omega + \Delta \omega) - \gamma^*(\omega)\) \(\leq \epsilon\).

Indeed,

1. Consider \(x, j\) and \(\gamma^*\) such that \(||\Delta x|| \leq \delta\) and \(||A_j||_p \leq \gamma^* ||x||_p\). Then, by a reasoning similar to the derivation of Proposition 3 and since by definition of Problem (5), \(\gamma^* \leq \max_{A \in \mathcal{A}^d} ||A||\), we have:

\[
||A_j(x + \Delta x)||_p \leq (\gamma^* + 2\max_{A \in \mathcal{A}^d} ||A|| ||\delta_k||_p) ||x + \Delta x||_p \leq (\gamma^* + \epsilon) ||x + \Delta x||_p
\]

by definition of \(\delta\).

2. Conversely, consider \(x, j\) and \(\gamma^*\) such that \(||\Delta x|| \leq \delta\) and \(||A_j(x + \Delta x)||_p \leq \gamma^* ||x + \Delta x||_p\). Then,

\[
||A_j(x + \Delta x)||_p \leq ||A_j||_p ||x + \Delta x||_p \leq (\gamma^* + \epsilon) ||x + \Delta x||_p \leq (\gamma^* + \epsilon) ||x||_p
\]

Hence, since \(\sqrt{x^2 - y^2} \leq \sqrt{||x - y||^2}\), the result holds. Finally, from the convergence of \(\omega_k\), for any \(0 < \epsilon < 1\), there exists \(l\) such that the components in \(\omega_k\) corresponding to the modes do not change (or have reached the limit) for all \(k > l\), which concludes the proof.

Appendix B. PROOF OF PROPOSITION 4

We define an essential set of a sampled problem as a subset, of minimal cardinality, of constraints whose removal implies a decrease in the cost of the problem. Problem (14) is convex and admits a unique solution. Moreover, there exists a sample \((x, j) \in \omega_N\) such that \(\lambda^*(\omega_N) = ||A_j||_p\), that is \((x, j) = \text{argmax}_j ||A_j||_p\), so that one is a bound on the cardinality of its essential sets. Let \(V(\omega_N)\) be the set \(\{(x, j) \in Z_t : ||A_j||_p > \lambda^*(\omega_N)\}\) of violated constraints. We seek to bound the probability to sample \(\omega_N\) such that the measure of this set is larger than \(\epsilon_1\). We consider two cases, depending on whether or not the essential set of Problem (14) is unique.

Suppose the essential set is unique with probability one. Then, a classical result from classical constrained optimization ([Calafiore, 2010, Thm. 3.3]) ensures that for all \(\epsilon_1 \in [0, 1]\), we have \(\mu^N_{\omega_N} \subset Z_t^N : \mu(V(\omega_N)) \leq \epsilon_1 \geq \beta_1(\epsilon_1, N)\), where \(\beta_1(\epsilon_1, N) = 1 - (1 - \epsilon_1)^N\).

Conversely, suppose that the problem admits distinct essential sets with probability larger than 0, which implies the existence of at least one matrix \(A \in \mathcal{A}^d\) such that \(A^TA = \gamma_1 I_d\) for some \(\gamma \in \mathbb{R}_{>0}\) (adapted from [Berger et al., 2021]). Let \(\lambda^* = ||A||_p\) be the solution to Problem (13). We consider the two following cases:

(a) \(A^TA = \lambda^* I_d\). Then, \(\mu(V(\omega_N))\) can only be positive if \(\omega_N\) contains no \((x, j)\) such that \(A_j = \text{A}_j\), which happens with probability \((1 - \frac{1}{m})^N\). Hence, \(\forall \epsilon_1 \in [0, 1]\),

\[
\mu_1^N_{\omega_N} \subset Z_t^N : \mu(V(\omega_N)) \geq \epsilon_1 \leq 1 - \frac{1}{m}^N \leq (1 - \epsilon_1)^N.
\]

(b) \(A^TA \neq \lambda^* I_d\). Let \(V(\lambda) = \mu\{(x, j) \in Z_t : ||A_j||_p > \lambda\}\). Then, since \(1 - V(\lambda)\) is the cumulative distribution function of \(||A_j||_p\), \(V(\lambda)\) is right continuous and non-increasing. Hence, \(\forall \epsilon_1 \in (0, 1)\), there exists an \(\lambda_1(\epsilon_1) \in \mathbb{R}_{>0}\) satisfying \(V(\lambda_1(\epsilon_1)) = \epsilon_1\), where \(\lambda_1(\epsilon_1)\) is a decreasing function of \(\epsilon_1\). In addition, \(\lambda^*(\omega_N) \leq \lambda_1(\epsilon_1)\) implies that \(\omega_N\) contains no \((x, j)\) in \(\{(x, j) \in Z_t : ||A_j||_p > \lambda_1(\epsilon_1)\}\), which occurs with probability \((1 - \epsilon_1)^N\).

Hence,

\[
\mu_1^N_{\omega_N} \subset Z_t^N : \mu(V(\omega_N)) \geq \epsilon_1 \leq 1 - (1 - \epsilon_1)^N.
\]

In all cases, \(\mu_1^N_{\omega_N} \subset Z_t^N : \mu(V(\omega_N)) \leq \epsilon_1 \geq \beta_1(\epsilon_1, N)\), where \(\beta_1(\epsilon_1, N) = 1 - (1 - \epsilon_1)^N\). This result can replace Theorem 4 in [Kenan et al. 2019]. Then, the proof follows the same lines as the derivation of Theorem 6 in [Kenan et al. 2019], so that details are omitted here.

Appendix C. PROOF OF THEOREM 2

Throughout this section, let \(D\) be given by (17). Before proceeding to the proof itself, some additional definitions and properties are needed.

**Proposition 7.** Given \(x, y \in \mathbb{R}^n\), the \(k\)-th Kronecker power of \(x\) is defined as: \(x^\otimes k = x \otimes x^{(k-1)}\) and \(x^\otimes 1 = x\), where \(\otimes\) denotes the Kronecker product. Then, \(x^{[d]} = C_d x \otimes x^{[d-1]}\), where \(C_d \in \mathbb{R}^{D \times n^d}\) is a matrix of norm \(||C_d|| \leq \sqrt{D}\) and \(x^{[d]}\) is the d-lift of \(x\).

Moreover, the following relations hold:

\[
||x^{\otimes d}|| = ||x||^k
\]
\[(x + y)^{\otimes d} = (x + y) \otimes \ldots \otimes (x + y) = \sum_{k=0}^{d} K_{xy}(d, k)\]
\[= x^{\otimes d} + y^{\otimes d} + \sum_{k=1}^{d-1} K_{xy}(d, k)\]  \hspace{1cm} \text{(C.2)}

where \(K_{xy}(d, k)\) is the sum of all possible sequences of \(d\) Kronecker products composed of \(d - k\) copies of \(x\) and \(k\) copies of \(y\). For \(k = 0, \ldots, d\), \(K_{xy}(d, k)\) is composed of \(\binom{d}{k}\) terms of norm \(\|x^{\otimes(d-k)} \otimes y^{\otimes k}\| = \|x\|^{(d-k)}\|y\|^k\).

**Proof of Proposition 7:** The two last relations are well-known. In addition, \(x^{\otimes d}\) is composed of \(d^\alpha\) copies of each monomial of degree \(d\) \(x^\alpha\). Hence, if the only non-zero entries of \(\alpha\)-th row of \(C_d\) are \(1\) and lie at columns associated to monomials \(x^\alpha\), \(C_x^{\otimes d} = x^d\). Moreover, \(\|C_d\|^2 \leq \|C_d\|_F^2 = \sum_\alpha \alpha! \frac{1}{\alpha!} = D\).

We are now ready to prove Theorem 2.

**Proof.** Let \(\gamma = \gamma^*_{sos}(\omega_N)\) and \(P = P^*_{sos}(\omega_N)\). By definition,
\[\|([A_j^d]x^d)\|_P \leq \gamma^d\|x^d\|_P, \forall (x, j) \in \omega_N.\]  \hspace{1cm} \text{(C.3)}

Let \(\psi = \|(x + \Delta x)^d\|_P\). For any \(\Delta x\) such that \(\|\Delta x\| \leq \Delta(e)\) and \(x + \Delta x \in S\), we obtain that \(\forall (x, j) \in \omega_N\)
\[\|([A_j^d](x + \Delta x))\|_P \leq \|([A_j^d]C_d(x + \Delta x)^{\otimes d})\|_P\]  \hspace{1cm} \text{(C.2)}
\[\leq \|([A_j^d]C_d x^{\otimes d})\|_P + \|([A_j^d]C_d(\Delta x)^{\otimes d})\|_P\]  \hspace{1cm} \text{(C.3)}
\[\leq \gamma^d\|x^d\|_P + \|([A_j^d]C_d\Delta x^{\otimes d})\|_P\]  \hspace{1cm} \text{(10)}
\[\leq \gamma^d\|([A_j^d]\Delta x^{\otimes d})\|_P + \|([A_j^d]C_d\Delta x^{\otimes d})\|_P\]  \hspace{1cm} \text{(19)}
\[\leq \gamma^d\psi + (\gamma^d + \|A_j^d\|d)\|\Delta x^d\|_P\]  \hspace{1cm} \text{(11)}
\[\leq \left(\gamma^d + B(M^d)\Delta(e)^d\kappa(P)\right)\psi + \left(\gamma^d + B(M^d)\sqrt{D}((1 + \Delta(e))^d - 1 - \Delta(e)^d)\kappa(P)\right)\psi\]  \hspace{1cm} \text{(11)}
\[\leq \left(\gamma^d + (\gamma^d + B(M^d))\Delta(e)^d\kappa(P)\right)\psi + (\gamma^d + B(M^d)\sqrt{D}\sum_{k=0}^{d} \binom{d}{k}\Delta(e)^k - 1 - \Delta(e)^d))\kappa(P)\psi\]  \hspace{1cm} \text{(10)}

Let \(\epsilon \in [0, 1]\), \(\Delta(e) = \\sqrt{2 - 2\epsilon}\) and \(\epsilon_{sos}(\beta, m, N) = m^\epsilon \left(1 - \frac{2(1 - \frac{1}{m})}{m\epsilon}\right)\). The extended version of Proposition 2 guarantees the existence of a set \(\omega'_N = \{x_i + \Delta x_i, j_i\}, i = 1, \ldots, N\) such that for \(i = 1, \ldots, N, \Delta x_i \leq \Delta(e)\) and \(x_i + \Delta x_i \in S\) and such that \(\gamma^*_{sos}(\omega'_N) = \gamma^*_{sos}(\omega_N)\) with probability at least \(\beta\). Hence, \(\|([A_j^d]x^d)\|_P \leq \gamma^d + ((\gamma^d + B(M^d))\sqrt{D})\|([A_j^d]x^d)\|_P\), \(\forall (x, j) \in \omega'_N\). Following the same reasoning as in Theorem 1 concludes the proof.