Effects of short-term ambient PM$_{2.5}$ exposure on cardiovascular disease incidence and mortality among U.S. hemodialysis patients: a retrospective cohort study

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Abstract

Background: Ambient PM$_{2.5}$ is a ubiquitous air pollutant with demonstrated adverse health impacts in population. Hemodialysis patients are a highly vulnerable population and may be particularly susceptible to the effects of PM$_{2.5}$ exposure. This study examines associations between short-term PM$_{2.5}$ exposure and cardiovascular disease (CVD) and mortality among patients receiving maintenance in-center hemodialysis.

Methods: Using the United State Renal Data System (USRDS) registry, we enumerated a cohort of all US adult kidney failure patients who initiated in-center hemodialysis between 1/1/2011 and 12/31/2016. Daily ambient PM$_{2.5}$ exposure estimates were assigned to cohort members based on the ZIP code of the dialysis clinic. CVD incidence and mortality were ascertained through 2016 based on USRDS records. Discrete time hazards regression was used to estimate the association between lagged PM$_{2.5}$ exposure and CVD incidence, CVD-specific mortality, and all-cause mortality adjusting for temperature, humidity, day of the week, season, age at baseline, race, employment status, and geographic region. Effect measure modification was assessed for age, sex, race, and comorbidities.

Results: Among 314,079 hemodialysis patients, a 10 µg/m$_3$ increase in the average lag 0–1 daily PM$_{2.5}$ exposure was associated with CVD incidence (HR: 1.03 (95% CI: 1.02, 1.04)), CVD mortality (1.05 (95% CI: 1.03, 1.08)), and all-cause mortality (1.04 (95% CI: 1.03, 1.06)). The association was larger for people who initiated dialysis at an older age, while minimal evidence of effect modification was observed across levels of sex, race, or baseline comorbidities.

Conclusions: Short-term ambient PM$_{2.5}$ exposure was positively associated with incident CVD events and mortality among patients receiving in-center hemodialysis. Older patients appeared to be more susceptible to PM$_{2.5}$-associated CVD events than younger hemodialysis patients.

Keywords: Hemodialysis patients, Susceptible population, Air pollution, PM$_{2.5}$, Short-term exposure

Background

Short-term exposure to PM$_{2.5}$ (particles with a mass median aerodynamic diameter of less than 2.5 µm) is an established cause of cardiovascular disease and mortality [1], and is a criteria air pollutant regulated by US National Ambient Air Quality Standards (NAAQS). When inhaled, PM$_{2.5}$ particles can travel deep into the
lung and even enter the circulatory system [2]. PM$_{2.5}$ particles can trigger inflammation (both systemic and local) and oxidative stress, accelerate atherosclerosis, and alter cardiac autonomic function which can eventually lead to cardiovascular diseases and ultimately deaths [3—8].

Previous studies have established ambient PM$_{2.5}$ as a cause of cardiovascular disease and mortality in the general population; however, less is known about the risks associated with PM$_{2.5}$ exposure among subgroups of the population that may be particularly susceptible to PM$_{2.5}$-associated health effects [1, 9]. One such group are patients with kidney failure requiring dialysis. These individuals have a higher cardiovascular disease (CVD) risk than the population, and US dialysis patients have an annual all-cause mortality risk of 16% [9—13]. Quantifying the mortality risk from air pollution is a critical step in identifying modifiable adverse health outcome risk factors in this population. Yet, to our knowledge, there has been no prior investigation of the short-term effects of ambient PM$_{2.5}$ on CVD incidence or mortality in this patient population.

Therefore, we assessed the association of short-term ambient PM$_{2.5}$ exposure with new CVD events, and with mortality among patients receiving maintenance in-center hemodialysis (HD) in the contiguous United States, considering lagged effects of exposure (up to 3 days after exposure) and potential effect measure modifications by age, sex, race, and baseline comorbidities.

Methods

Study setting and study population

We identified a retrospective cohort of maintenance hemodialysis patients from the United States Renal Data System (USRDS), a claim-based national registry of renal disease patients. USRDS includes nearly all dialysis patients in the United States with individual level data on patients’ demographics, hospitalizations, and provider information [14]. Baseline demographic characteristics, dialysis treatment, and comorbidity status were extracted. We included patients aged 18 years or older who initiated in-center maintenance hemodialysis treatment between Jan 1st 2011 and Dec 31st 2016 in the contiguous U.S. who had Medicare as primary payer. Participants were followed from the fourth month of dialysis initiation until first incident diagnosis of CVD, change of renal disease treatment modality (e.g. receiving kidney transplant), loss of Medicare coverage, death, or administrative end of study follow-up (Dec 31st 2016).

Exposure assessment and weather data

For each participant-day under follow-up, a daily record of location, at the ZIP code-level, was assigned based on the ZIP code of the dialysis clinic visited during that period. Patients typically visited dialysis clinics every 2 to 3 days. For a date between claim reporting periods (with no information on clinic visits), the ZIP code of the last clinic visited was assigned (1.3% of the person-days in the analysis had a ZIP code that was different from the previous person-day of the same individual). Previous studies have shown that patients in the USRDS cohort have a median travel distance of 5.7 miles to their dialysis clinic [15, 16]. Thus, ZIP code assignment based on dialysis clinic location should be a relatively accurate surrogate for hemodialysis patients’ geographic location of residence during study follow-up.

ZIP code-level daily PM$_{2.5}$ concentrations were estimated with a previously described prediction model [17, 18]. In brief, this model estimates daily average PM$_{2.5}$ on a 1×1 km grid for the entire contiguous U.S. by incorporating multiple inputs like chemical transport model simulations, satellite aerosol optical depth measurements, meteorology, and land use. Gridded PM$_{2.5}$ estimates were then converted to population-weighted ZIP-level estimates using 2010 Census tract population values. Daily temperature and relative humidity data were estimated based on records from the U.S. National Oceanic and Atmospheric Administration (NOAA). In detail, daily temperature and relative humidity from weather station measurements were interpolated to census tract centroids and then averaged over tracts to ZIP codes. A patient’s daily record of location was linked to an estimate of the daily concentration of PM$_{2.5}$ at that location, as well as an estimate of the temperature and relative humidity for that date and ZIP code.

Outcome assessment

Follow-up to ascertain incident CVD events, CVD-related deaths, and vital status through December 31, 2016 was based on USRDS records [19]. The outcomes of interest were all-cause mortality, CVD-specific mortality, and incident CVD event. The primary cause of death recorded in ESRD Death Notification Form (CMS 2746) was used to classify mortality into all-cause and CVD-specific mortality. We defined incident CVD events as the first CVD related emergency department (ED) visit, hospitalization, or mortality, which ever occurred first since the follow-up start. It is worth noting that, conventionally, “incident CVD event” or “CVD incidence” refers to those first events occurring among participants determined to be free of CVD at baseline only, and in this analysis, these terms’ meaning is different from their conventional definition by expanding to all participants regardless of their baseline CVD status. International Classification of Disease, 9th and 10th Revision (ICD-9, ICD-10) codes were used to identify cause-specific ED visit and inpatients cardiovascular events. In brief, any
ED and hospitalization with a code indicating cerebrovascular, ischemic heart disease, hypertensive disease, heart failure, cardiac arrest, and related incidence were classified as CVD related ED and hospitalization events. The cause of death codes and International Classification of Disease codes used are summarized in Supplemental Table 1.

Statistical analysis

We utilized discrete time hazards models to assess the association between ZIP code-level short-term PM$_{2.5}$ exposure with mortality and CVD events [20]. Discrete time hazards models have been widely used in occupational and environmental cohort studies [21–23]. This approach is similar to a Poisson regression model applied to a person-period data structure, which allows controlling for individual level factors [20]. The equation for the fully-adjusted model is presented in Eq. 1, where $Y$ is the individual specific response at time $t$ (e.g. 0 if no event, 1 event,...) and $i$, $t$, and $k$ index individual, time (day), and number of lag days respectively. We used a generalized linear model $E(Y_{it}|\text{covariates}) = g^{-1}(\eta_{it})$ with a complementary log–log link function $g$ and a linear predictor $\eta_{it}$ as follows:

$$
\log(-\log(1 - E(Y_{it} | \text{covariates}))) = \eta_{it} = \text{factor(MonthSinceDialysis)}_{i} + \sum_{k=0}^{K} \beta_{1k} PM2.5_{itk} + \beta_{2} Temperature_{it} + \beta_{3} RelativeHumidity_{it} + \beta_{4} factor(DOW)_{it} + \beta_{5} factor(Season)_{it} + \beta_{6} Age(DialysisInitiation)_{i} + \beta_{7} factor(Race)_{i} + \beta_{8} factor(EmploymentStatus)_{i} + \beta_{9} factor(Region)_{i}.
$$

Variables DOW, Season, Race, Region and Employment status were categorical thus the coefficients represent vectors rather than scalars. Hazard ratios (HRs) per 10-μg/m³ increase in PM$_{2.5}$, together with corresponding 95% confidence intervals (95% CIs) were estimated. We incorporated time since initiation of dialysis in the model as the analysis time scale to control for the potential temporal variation in baseline risk of the outcome of interest by this factor, which has been suggested to be an important predictor of mortality among dialysis patients in prior studies [24, 25]. Time-varying confounders including: daily temperature, daily relative humidity, day of the week (DOW), and season (warm vs. cold) were controlled for to adjust for potential time varying confounding. Specifically, DOW and season were incorporated to account for short- and long-term temporal variability of the outcomes [26, 27]. We tested non-linear functional forms of temperature and relative humidity in the model and found no significant change in effect estimates and model fitness when compared to linear forms. Thus, we modeled temperature and relative humidity as linear variables in the final model. Time-invariant confounders included age at dialysis initiation, race, baseline employment status, and geographic region (Northeast, Midwest, South, and West). Baseline employment status (employed, unemployed, retired, other) was incorporated as a surrogate for socio-economic status (SES).

Short-term PM$_{2.5}$ exposure effects were assessed for the same day, lag days, and moving averages across different lag periods to assess the persistent effect of PM2.5 exposure across multiple days. The same day (lag day 0) effects were assessed in a model for that same day PM$_{2.5}$ and its association with the outcome. The lagged effects of prior days’ exposures were assessed in a model that associated lag day 0, 1, 2, and 3 exposures with the outcome in the same model. The average lagged effects of prior exposure days were assessed in a model that associated the average of daily PM$_{2.5}$ across lag periods 0–1, 0–2, or 0–3 days prior. For example, the moving average of lag period 0–1 was calculated by averaging the exposure levels of the day of the event and the day preceding the event, which is later incorporated in the model as the sole exposure variable. Among different models assessing short-term lagged effect, the model with the lower Akaike information criterion (AIC) value was considered to be the better model depicting the short-term effects.

Potential effect measure modification by baseline comorbidity status, race, sex, and age at dialysis initiation were assessed. The comorbidities assessed were diabetes, CVD, and chronic obstructive pulmonary disease (COPD) reported at the start of maintenance hemodialysis on CMS 2728 form. Specifically, the patients who reported to have any of the following conditions: atherosclerotic heart disease (ASHD), cardiac arrest, congestive heart failure (CHF), cerebrovascular disease, cardiac dysrhythmia, ischemic heart disease (IHD), myocardial infarction (MI), pericarditis, peripheral vascular disease, and other cardiac disease, were classified to have CVD at baseline. The patients who reported to have diabetes (primary or contributing), who were diabetic and currently on insulin, on oral medications, or without medications, or with diabetic retinopathy were classified to have diabetes at baseline. The patients who reported to have chronic obstructive pulmonary disease were classified to have COPD at baseline.

Likelihood ratio tests for homogeneity were performed by comparing models with and without an interaction term between the potential modifier and PM term. A small Chi-square test of likelihood ratios test p-value (less than 0.05) suggested the rejection of effect homogeneity.
hypothesis, which indicates PM effects are not similar across different levels of potential modifiers.

The data management and statistical analysis were performed with SAS software, Version 9.4 (SAS Institute Inc., Cary, NC, USA). Plots were generated with R Project for Statistical Computing Version 3.5.0 (R Foundation for Statistical Computing, Vienna, Austria). The study protocol was reviewed and exempted by the Institutional Review Board of the University of North Carolina on Nov 16, 2020 (Study #:20–1469).

Results
The final cohort included a total of 314,079 hemodialysis patients with 193,121,928 person-days of follow-up in the mortality analysis and 94,385,773 person-days in the CVD event analysis. Between 2011 and 2016, there were 82,567 deaths, 35,857 (43.4%) of which were CVD-related mortalities, and 208,113 patients experienced at least one CVD event.

Table 1 displays the study population characteristics. At the initiation of hemodialysis, more than half of the participants were older than 65 years of age. The average vintage period, which is the time between dialysis initiation and the end of follow-up, was 1.9 (SD=1.4) years. Approximately 44% of the study participants were female, and 65% were white. At hemodialysis initiation, more than half of the cohort was reported to have diabetes, about 60% was reported to have CVD, and 10% was reported to have COPD. Patients who died during the study period were likely to be older at dialysis initiation compared to the overall study population or the cohort who experienced at least one CVD event.

The distribution of ZIP code-level PM$_{2.5}$ concentration is summarized in Supplemental Table 2. For the mortality analysis, the average daily ZIP code-level PM$_{2.5}$ was around 8.5 (±4.8) µg/m$^3$ with an inter-quartile range (IQR) of 5.2 µg/m$^3$. For CVD incidence analysis, the PM$_{2.5}$ exposure distribution was similar to that of the mortality analysis with a mean of 8.6 (±4.9) µg/m$^3$ and IQR of 5.3 µg/m$^3$.

Exposure to short-term daily ambient PM$_{2.5}$ was associated with an elevated hazard of all-cause mortality (Fig. 1, Supplemental Table 3). The HR associated with a 10-µg/m$^3$ increase in same day, lag-day 1, lag-day 2, and lag-day 3 PM$_{2.5}$ were 1.01 (95% CI: 0.99, 1.03), 1.03 (95% CI: 1.01, 1.06), 1.00 (95% CI: 0.97, 1.02), and 1.00 (95% CI: 0.98, 1.02), respectively (Fig. 1, Supplemental Table 3). The short-term PM effects were most prominent for the same day and lag-day 1 exposure compared to the overall study population or the cohort who experienced at least one CVD event.

Similar to the effect measure modification effect observed for all-cause mortality, HD patients who initiated dialysis in the oldest age group (≥75 years), were of white race, and did not have diabetes at dialysis initiation experienced relatively higher PM effects on CVD-specific mortality compared to other groups by these potential modifiers (Fig. 2). The highest HR of 1.07 (95% CI: 1.05, 1.10) associated with every 10-µg/m$^3$ increase in average daily PM$_{2.5}$ across lag-day 0 and 1 was observed among patients who initiated dialysis at the age of 75-year or older (Fig. 2). The likelihood ratio test for homogeneity also suggested PM mortality effects are not homogeneous across levels of dialysis initiating age, race, and baseline diabetes status with low p-values (Supplemental Table 4).

Exposure to short-term daily ambient PM$_{2.5}$ was associated with an elevated hazard of CVD-specific mortality (Fig. 1, Supplemental Table 3). The HR associated with a 10-µg/m$^3$ increase in same day, lag-day 1, lag-day 2, and lag-day 3 PM$_{2.5}$ were 1.02 (95% CI: 0.99, 1.05), 1.03 (95% CI: 0.99: 1.07), 0.99 (95% CI: 0.96, 1.03), and 1.01 (95% CI: 0.98, 1.04), respectively (Supplemental Table 3). The short-term PM effects were most prominent for the same day and lag-day 1 exposure for CVD-specific mortality. The HR associated with a 10-µg/m$^3$ of the average daily PM$_{2.5}$ across same-day and lag-day 1 was 1.05 (95% CI: 1.03, 1.07) associated with every 10-µg/m$^3$ increase in average daily PM$_{2.5}$ across same-day and lag-day 1 (Supplemental Table 3).

Exposure to short-term daily ambient PM$_{2.5}$ was associated with a higher hazard of incident CVD events (Fig. 1, Supplemental Table 3). The HR associated with a 10-µg/m$^3$ increase in same day, lag-day 1, lag-day 2, and lag-day 3 PM$_{2.5}$ were 1.02 (95% CI: 1.01, 1.03), 1.01 (95% CI: 1.00, 1.03), 1.00 (95% CI: 0.98, 1.01), and 1.00 (95% CI: 0.99,
| Characteristics                  | Total | Mortality | Incidence |
|---------------------------------|-------|-----------|-----------|
|                                 |       | All-Cause | CVD-Specific | Incident CVD Event |
| n                               | 314,079 | 82,567 | 35,857 | 208,113 |
| Person-day                      | 193,121,928 | 193,121,928 | 193,121,928 | 94,385,773 |
| Age at HD Initiation (year)     | 63.6 (14.6) | 69.0 (13.0) | 69.4 (13.2) | 65.4 (14.0) |
| Vintage (year)<sup>a</sup>      | 1.9 (1.4) | 1.7 (1.2) | 1.7 (1.2) | 2.2 (1.5) |
| Vintage (year)                  |       |           |           |           |
| Less than 1 year                | 108,957 (34.7) | 32,153 (38.9) | 13,596 (37.9) | 54,144 (26.0) |
| 1 – 2 years                     | 81,617 (26.0) | 23,347 (28.3) | 10,145 (28.3) | 55,742 (28.6) |
| 2 – 3 years                     | 53,198 (16.9) | 14,035 (17.0) | 6,266 (17.5) | 40,457 (20.4) |
| 3 – 4 years                     | 35,163 (11.2) | 8,058 (9.8) | 3,644 (10.2) | 28,315 (13.6) |
| Over 4 years                    | 35,144 (11.2) | 4,974 (6.0) | 2,206 (6.2) | 29,455 (14.2) |
| BMI (kg/m<sup>2</sup>)          |       |           |           |           |
| Female                          | 131,209 (43.9) | 36,343 (44.0) | 15,244 (42.5) | 93,832 (45.1) |
| Race                            |       |           |           |           |
| White                           | 194,790 (65.1) | 58,106 (70.4) | 25,103 (70.0) | 136,196 (65.4) |
| Black                           | 88,415 (29.6) | 20,071 (24.3) | 8,917 (24.9) | 60,123 (28.9) |
| Other                           | 15,868 (5.3) | 4,390 (5.3) | 1,837 (5.1) | 11,794 (5.7) |
| Reported Cause of Kidney Failure|       |           |           |           |
| Diabetes                        | 152,106 (48.4) | 40,205 (48.7) | 18,734 (52.3) | 105,544 (50.7) |
| Glomerulonephritis              | 20,846 (6.6) | 3,536 (4.3) | 1,368 (3.8) | 11,351 (5.5) |
| Hypertension                    | 97,204 (31.0) | 26,555 (32.2) | 11,530 (32.2) | 64,554 (31.0) |
| Other                           | 43,923 (14.0) | 12,271 (14.9) | 4,225 (11.8) | 26,662 (12.8) |
| Access Type                     |       |           |           |           |
| AVF                             | 54,008 (18.1) | 11,094 (13.4) | 4,937 (13.8) | 32,914 (15.8) |
| Graft                           | 9,527 (3.2) | 2,427 (2.9) | 1,070 (3.0) | 6,543 (3.1) |
| Catheter                        | 235,042 (78.6) | 66,320 (80.3) | 28,782 (80.3) | 159,809 (76.8) |
| Other                           | 496 (0.2) | 139 (0.2) | 61 (0.2) | 327 (0.2) |
| Baseline Comorbidities          |       |           |           |           |
| CVD                             | 160,760 (53.8) | 53,452 (64.7) | 24,218 (67.5) | 123,121 (59.2) |
| COPD                            | 30,417 (10.2) | 11,841 (14.3) | 5,220 (14.6) | 24,067 (11.6) |
| Diabetes Mellitus               | 178,489 (59.7) | 49,725 (60.2) | 22,728 (63.4) | 128,162 (61.6) |
| Geographic Region               |       |           |           |           |
| Northeast                       | 68,668 (21.9) | 20,183 (24.4) | 8,831 (24.6) | 47,707 (22.9) |
| Midwest                         | 59,243 (19.9) | 15,362 (18.6) | 6,033 (16.8) | 39,405 (18.9) |
| West                            | 40,001 (12.7) | 9,091 (11.0) | 4,146 (11.6) | 24,398 (11.7) |
| South                           | 146,167 (46.5) | 37,931 (45.9) | 16,847 (47.0) | 96,603 (46.4) |
| Baseline Employment Status      |       |           |           |           |
| Employed                        | 30,746 (9.8) | 4,621 (5.6) | 2,082 (5.8) | 15,655 (7.5) |
| Retired                         | 196,525 (62.6) | 61,159 (74.1) | 26,353 (73.5) | 139,909 (67.2) |
| Unemployed                      | 76,338 (24.3) | 15,088 (18.3) | 6,793 (18.9) | 47,033 (22.6) |
| Other                           | 10,470 (3.3) | 1,699 (2.1) | 629 (1.8) | 5,516 (2.7) |

<sup>a</sup>Vintage is defined as the time between hemodialysis initiation and the end of study follow-up, which is death, drop-out of Medicare, change of modality, or the last day of study period

<sup>HD</sup> Hemodialysis, BMI Body Mass Index, AVF Arteriovenous Fistula, COPD Chronic obstructive pulmonary disease
1.01), respectively (Supplemental Table 3). The short-term PM effects were most prominent for the same day and lag-day 1 exposure for CVD related events. Effects at lag-day 2 and lag-day 3 daily PM$_{2.5}$ were not observed (Supplemental Table 3). The HR associated with a 10-µg/m$^3$ of the average daily PM$_{2.5}$ across same-day and lag-day 1 was 1.03 (95% CI: 1.03, 1.04) (Fig. 1, Supplemental Table 3). The HRs associated with long lag periods (lag 0–2 or lag 0–3) average daily PM$_{2.5}$ was very similar to the HR of the lag 0–1 average PM$_{2.5}$ (Supplemental Table 3).

For CVD events, none of the factors assessed significantly modified the short-term PM effect. Slightly higher hazards were observed among hemodialysis patients who had COPD at baseline and who were males with overlapping confidence interval ranges (Fig. 4). The likelihood ratio test of homogeneity also indicates the PM$_{2.5}$ effects were similar across different levels of dialysis initiating age, race, and baseline diabetes status for incident CVD events with greater than 0.05 p-values (Supplemental Table 4).

Discussion

We estimated associations between short-term PM$_{2.5}$, all-cause mortality, CVD-specific mortality, and incident CVD among US hemodialysis patients during the period 2011–2016. Short-term PM$_{2.5}$ exposure was associated with elevated hazards for all-cause mortality, CVD-specific mortality, and incident CVD among the study population. Each 10 µg/m$^3$ increase in short-term PM$_{2.5}$ was associated with a 4%, 5%, and 3% increase in the hazard of all-cause mortality, CVD-specific mortality, and incident CVD events, respectively. We also observed relatively higher PM$_{2.5}$ hazards on all-cause mortality and CVD-specific mortality among hemodialysis patients who initiated dialysis at an older age, who were white, and who did not report having diabetes mellitus at the initiation of dialysis.

The effects of PM$_{2.5}$ exposure among US hemodialysis patients have not been extensively evaluated. To our knowledge, this is the first study to examine short-term ambient PM$_{2.5}$ exposure effects among US hemodialysis patients. We previously observed an elevated mortality risk associated with wildfire generated PM$_{2.5}$ exposure among maintenance hemodialysis patients [28]. In addition to mortality, we also observed an elevated CVD incidence hazard associated with short-term PM$_{2.5}$ exposure, which is consistent with a study among hemodialysis patients on the association between hospital admission and county-level short-term PM$_{2.5}$ for cardiovascular causes for the same-day PM$_{2.5}$ exposure [29]. Wyatt et al. reported a 0.9% (95% CI, 0.4, 1.8%) increase in hospital admission risk for cardiovascular causes for the same-day PM$_{2.5}$ exposure [29]. Recently, a study demonstrated an elevated risk of all-cause mortality associated with long-term (annual average) PM$_{2.5}$ among older (over 65-year old) US hemodialysis and peritoneal dialysis patients [30]. Their study reported that for very 10 µg/m$^3$ increase in annual average PM$_{2.5}$ exposure below 12 µg/m$^3$, there is a 1–5% increase in the HR of all-cause mortality, a similar range to our findings [30]. However, it was unexpected that the association between PM$_{2.5}$ and incident CVD was slightly lower than between PM$_{2.5}$ and CVD-specific mortality (3%
vs. 5% increase in HR per 10 µg/m$^3$ increase in PM$_{2.5}$, albeit with overlapping confidence intervals) because it is logical that a person would develop one or more CVD incidence(s) before he or she can die from it. However, this could be due to the different diagnosis coding regimes used to define CVD-specific mortality vs. CVD incidence. CVD mortalities were defined using CMS 2746 codes, which are not ICD-9 or -10 codes used for CVD incidence, and future studies should identify and use additional data sources on cause of death. Another possible explanation is that PM$_{2.5}$ exposure could impact the CVD case fatality rate, which could be assessed in future studies instead of first CVD event since follow-up to take all incident CVD events into account and to further understand PM’s impact on cardiovascular health among this population. In addition, evidence is emerging that elevated levels of fine particulate matter may contribute to incident chronic kidney disease, and its progression to severe kidney function impairment that requires maintenance dialysis treatment [31–33]. Together with previously published studies, our findings support the conclusion that there is an elevated health hazard associated with particulate matter exposure among dialysis patients. Furthermore, the daily ZIP code-level PM$_{2.5}$ for the majority (Q3 of 10.6 µg/m$^3$) of the included person-days in this analysis were below the current NAAQS regulatory standard of 35 µg/m$^3$ (24-h, 98th percentile). On average, PM$_{2.5}$ has been decreasing nationally in the US over the last decade [34]. However, the PM concentration in many highly-populated urban
areas still exceed the NAAQS, and extreme events like wildfires also lead to short-term high-level exposures. Our findings of elevated hazards of mortality and CVD conditions suggest that reduction of individual-level exposure may be beneficial to at risk individuals such as dialysis patients, a group that continues to increase in size in the US.

Our findings suggest that there is larger magnitude of association between \( \text{PM}_{2.5} \) and mortality (both all-cause and CVD-specific) among patients who initiated dialysis treatment at an older age compared to younger dialysis initiators (Figs. 2 and 3). The highest \( \text{PM}_{2.5} \) effects were observed among the oldest age group (75 or older at the initiation of dialysis). This effect modification observed could be explained by the higher prevalence of pre-existing conditions and higher mortality risk among this age group [35, 36]. To our knowledge, only one previous study has assessed the effect of PM exposure on outcomes by age groups among dialysis patients. Feng et al. observed a higher long-term \( \text{PM}_{2.5} \) exposure effect on all-cause mortality among dialysis patients who aged 75 years or over at dialysis initiation compared to those who initiated dialysis 65–74 years [30]. Even though Feng et al’s study population was different from our analysis, and they included peritoneal dialysis patients, our findings agrees that kidney disease patients who initiated dialysis at the age of 75 or older were likely to be particularly susceptible to \( \text{PM}_{2.5} \) exposure for all-cause mortality. This finding adds to the evidence that older hemodialysis patients should be targeted for exposure mitigation to reduce the overall public health burden in association with PM exposure.
In addition to older dialysis initiators, the results also suggest a larger magnitude of association between PM$_{2.5}$ and mortality among white patients than non-white patients. This difference association by race could be partially explained by differences in age distribution of dialysis patients by race (Supplemental Table 5). About 29.5% (15.8% in black, and 23.9% in other race group) of white patients initiated hemodialysis at age 75 years or older. Thus, it is likely that the higher PM effect observed among white patients could be driven by the larger proportion of white patients that initiated dialysis at an older age. Feng et al. reported a slightly higher long-term mortality effect among white patients for the area of annual PM$_{2.5}$ less than 12 µg/m$^3$ but a lower effect among white patients when restricted the analysis to area of annual PM$_{2.5}$ over 12 µg/m$^3$ [30]. Studies in other populations have also reported mixed results by race stratification [1]. In addition, this racial mortality hazards variation could be explained by the fact that the baseline mortality risks are different across race groups for ESKD patients. In a large cohort study on US dialysis patients, Yan et al. observed a consistently higher mortality risk among non-Hispanic whites in comparison to Hispanics and non-Hispanic blacks regardless of age [35]. Thus, it is not conclusive to determine how race is a modifier for PM$_{2.5}$ shot-term mortality effect, and more studies are needed to determine whether and how race modifies the association for dialysis patients as a social factor.

We also observed an un-expected higher hazard among patients with no diabetes at baseline. This difference in all-cause and CVD-specific mortalities hazards by baseline diabetes status could be partially explained by the
fact that more patients with diabetes (50.4% vs. 46.5%) initiated dialysis at a younger age (before 65-year old) (Supplemental Table 5). To our knowledge, no study has assessed the short-term particulate matter exposure effect by baseline diabetes status among dialysis patients. Feng et al.'s study on long-term PM effect, observed a higher risk among patients who reported diabetes as primary reason of end-stage kidney disease, which is different from baseline diabetes status [30]. Furthermore, studies among general population yields mixed and non-significant results regarding the potential modification by diabetes status on PM$_{2.5}$ effect [1, 37–39]. Thus, more studies are needed to determine comorbidity specific effect for fine particle exposure among dialysis patients.

This study has several strengths. First, the USRDS database is highly representative of hemodialysis patients in the US, with a nearly complete inclusion of those who initiated in-center hemodialysis between 2011 and 2016 in the US. The USRDS registry provides a complete registry of all deaths and healthcare service usages and contains detailed information regarding demographics, comorbidities, and dialysis clinic information. Second, linkage of the exposure using the dialysis treatment date and ZIP code of dialysis clinic visited enabled the most accurate exposure classification we can achieve with the available data. In our study population, more than half of the included patients had at least one ZIP code-level change of location during follow-up. This suggests that taking location changes into account for exposure linkage is necessary to avoid exposure misclassification for this population. Third, we utilized a sophisticated model of high spatial and temporal resolution with established accuracy to estimate the ZIP code-level daily PM$_{2.5}$ concentrations for the US [18, 40]. Last, the use of discrete time hazards model allowed for control for individual level time invariant factors together with time-varying factors to better estimate the health risk associated with fluctuation in ambient PM$_{2.5}$. This study design also allowed us to assess effect measure modification by individual-level factors (e.g. race, age).

This study also has some limitations. First, exposure misclassification is always a concern in environmental exposure studies as the individual-level exposure is not known. Instead, we assumed that the average exposure is similar across patients within in the same ZIP code area and that the degree of exposure misclassification does not vary simultaneously with daily variation in ZIP code-level exposure and health risks. Second, in this analysis we used information on employment at dialysis initiation to classify socio-economic status, which could result in some misclassification of SES. Employment at dialysis initiation may not necessarily reflect the patients’ SES. However, this is the best information available in our source data. Future studies should identify additional source of information or collect information for individual SES classification (e.g. income, education) among dialysis patients to better understand the potential confounding or modification effect by this factor. Third, information on additional potential effect measure modifiers was not available. Effect measure modification assessment for other relevant health risk factors such as blood pressure, dialysis adequacy, rest diuresis, liver diseases, dyslipidemia, cancer, vascular access type, and CV risk score should be performed to understand what other factors may modify the association between PM exposure and health outcomes. Finally, for all observational studies there is the possibility of unmeasured confounding (e.g. individual physical activity). In the presence of unmeasured confounding, our estimates could not be interpreted as causal.

Conclusions
In conclusion, in this large retrospective open cohort study, we found evidence of short-term effects of ambient PM$_{2.5}$ exposure on mortality and CVD outcomes among patients receiving maintenance in-center HD. For every 10 unit increase in ambient PM$_{2.5}$ the mortality and CVD hazards increased by 3–5%, and the mortality hazards associated with PM$_{2.5}$ was most prominent in older patients with a 10% increase in CVD-specific mortality hazard. These findings suggest that it is necessary for more studies to develop and implement interventions to manage particulate matter exposure in this and other particularly fragile populations.

Abbreviations
ASHD: Atherosclerotic heart disease; AVF: Arteriovenous fistula; BMI: Body mass index; CI: Confidence interval; CMS: Centers for Medicare & Medicaid Services; COPD: Chronic obstructive pulmonary disease; CVD: Cardiovascular disease; DOW: Day of the week; ED: Emergency department; EPA: Environmental Protection Agency; ESRD: End-stage renal disease; GEE: Generalized estimating equation; HD: Hemodialysis; HF: Heart failure; HR: Hazards ratio; ICD-10: International Classification of Disease 10th Revision; ICD-9: International Classification of Disease 9th Revision; IHD: Ischemic heart disease; IQR: Inter-quartile range; LRT: Likelihood ratio test; NAAQS: National Ambient Air Quality Standards; NOAA: National Oceanic and Atmospheric Administration; PM10: Integrated Science Assessment for Particulate Matter; PM2.5: Particulate matter with aero-diameter of 2.5 µm or smaller; SES: Socio-economic status; STD: Standard deviation; US: United States; USRDS: United States Renal Data System; WHO: World Health Organization.

Supplementary Information
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Additional file 1: Supplemental Table 1. Cause of death and ICD codes used to identify CVD incidences. Supplemental Table 2. United States 2011-2016 ZIP code-level daily ambient PM$_{2.5}$ concentration and meteorological data for all person-days included. Supplemental Table 3. Hazards
ratios (95% confidence interval) estimates per 10 µg/m$^3$ increase for the association between short-term PM$_{2.5}$ exposures and all-cause mortality, CVD-specific mortality, and incident CVD events among the overall study population. **Supplemental Table 4.** LRT Test For Homogeneity Results. **Supplemental Table 5.** Study cohort age at dialysis initiation distribution (%) by potential modifiers.

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**Authors’ contributions**
All authors contributed to the study design. YX, DBR, and AGR carried out the analysis. YX created the tables and plots and drafted the manuscript. All authors contributed to the interpretation of data, review, and editing. The authors read and approved the final manuscript.

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**Availability of data and materials**
The datasets sued in this analysis are available from the corresponding author and United States Renal Data System (USRDS) upon request.

**Declarations**

**Ethics approval and consent to participate**
The study protocol was reviewed and exempted by the Institutional Review Board of the University of North Carolina at Chapel Hill on Nov 16, 2020 (Study #20-1469).

**Consent for publication**
All authors approved the manuscript and this submission.

**Competing interests**
The authors declare that they have no competing interests.

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