Mapping dark matter with cosmic magnification

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We develop a new tool to generate statistically precise dark matter maps from the cosmic magnification of galaxies with distance estimates. We show how to overcome the intrinsic clustering problem using the slope of the luminosity function, because magnificability changes strongly over the luminosity function, while intrinsic clustering only changes weakly. This may allow precision cosmology beyond most current systematic limitations. SKA is able to reconstruct projected matter density map at smoothing scale $\sim 10''$ with S/N $> 1$, at the rate of 200-4000 deg² per year, depending on the abundance and evolution of 21cm emitting galaxies. This power of mapping dark matter is comparable to, or even better than that of cosmic shear from deep optical surveys or 21cm surveys.

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Introduction.— The precision mapping of the universe, and the accurate determination of cosmological parameters have been enabled by the recent generation of cosmic microwave background (CMB) experiments, galaxy and lensing surveys, and new analysis techniques. Weak gravitational lensing has emerged with a promising future of mapping dark matter directly, which would allow the inference of the state of the universe, including its dynamics and the nature of dark energy. Lensing is free from modeling assumptions, and can be accurately predicted from first principles. Several major surveys are underway, under construction or in the planning stage. Currently, most attention has focused on using the lensing induced cosmic shear. But such an approach is subject to a series of difficult experimental systematics and 21cm background lensing are promising. But contaminations such as the kinetic Sunyaev Zeldovich effect and/or non-Gaussianity may degrade their accuracy. In this paper we will address an alternative approach, the lensing induced cosmic magnification, which is not subject to the known problems, and could provide a robust statistical signal.

Traditionally, intrinsic clustering had presented a serious problem to measurement of cosmic magnification. The observable quantity is the surface density of galaxies above some flux threshold. A variation in this surface density is then interpreted as lensing. Unfortunately, intrinsic clustering is usually larger than the lensing induced signal. By utilizing the redshift information, intrinsic clustering can be effectively eliminated in lensing correlation functions. In this paper, we further show that, beyond the above statistical lensing measurement, 2D convergence $\kappa$ maps can be reconstructed with lower systematics and larger sky coverage than cosmic shear maps, by utilizing both the redshift and flux information of galaxies. 2D $\kappa$ maps not only provide independent and robust constraints on cosmology, but also are complementary to traditional shear maps. It allows one to explicitly and locally solve for non-reduced shear, an independent mode of checking E-B decomposition, and break the mass-sheet degeneracy.

Cosmic magnification.— Cosmic magnification causes coherent changes in the apparent galaxy number density. Let $N_{ij}$ be the observed number of galaxies (including false peaks) at the $i$-th flux bin and $j$-th redshift bin, falling into an angular pixel centered at direction $\hat{n}$ with angular size $\theta$. It can be expressed as

$$N_{ij}(\hat{n}) = N_{ij} + N_{ij}^\alpha [W_{ij} \kappa_j(\hat{n}) + \delta_{ij}(\hat{n})] + \delta N_{ij}(\hat{n}). \quad (1)$$

The signal $W\kappa$ has unique dependence on galaxy flux through $W = 2(\alpha - 1)$. Here, $\alpha = -d\ln [dn/dF]/d\ln F - 1$ and $dn/dF$ is the mean number of observed galaxies per flux interval. $N_{ij} = N_{ij}^\alpha + N_{ij}^\delta$, $N_{ij}^\alpha$, $N_{ij}^\delta$, $N_{ij}^\delta$ are the mean number of detections, real galaxies and false peaks, respectively. $\delta_g$ and $\delta N_P$ are galaxy intrinsic clustering and Poisson fluctuation, respectively.

Our goal is to recover $\kappa$ of each angular pixel, given observables $N_{ij}$, $N_{ij}^\alpha$, $W_{ij}$ and $N_{ij}^\delta$. We consider SKA, which can detect $\sim 10^8$ high-z galaxies through the neutral hydrogen 21cm emission line. $\kappa$ has typical value $\sim 1\%$. To beat down Poisson fluctuations, $\gtrsim 10^4$ galaxies per angular pixel are required. Traditionally, objects are selected at a $5\sigma$ cut, where one can neglect the fraction of false detections. This of course also discards the majority of the signal. With a $0.5\sigma$ cut, one can reduce Poisson noise at $\theta \sim 10''$. To increase lensing signal while reducing $\delta_g$ contamination, we focus on source redshifts $z \gtrsim 2$. After averaging over the full redshift range $z \geq 2$, $\delta_g$ is still several times larger than $\kappa$. However, $\delta N_P$ and $\delta_{g,ij}$ have different flux dependence to that of the signal. Weighting each galaxies by some function of their flux can suppress the prefactors of $\delta_g$ and $\delta N_P$. Intuitively, Eq. 1 implies the optimal estimator to be linear in $N_{ij}$.

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The predictions rely on the assumed HI mass function $n(M_{\text{HI}}, z)$. We extrapolate the locally observed $n(M_{\text{HI}}, z) = n_0(z)(M_{\text{HI}}/M_\odot)^{-1.2}\exp(-M_{\text{HI}}/M_\odot)$ to high redshifts either assuming no evolution in both $n_0$ and $M_\star$ (conservative case) or $n_0(z), M_\star(z) \propto (1+z)^{1.45}\exp(-z/2.6)$ (realistic case), which is calibrated against Lyman-α observations (refer to [7] for details).

We adopt a flat ΛCDM cosmology with $\Omega_m = 0.3$, $\Omega_\Lambda = 0.7$, $h = 0.7$, $\sigma_8 = 0.9$, the primordial power index $n = 1$, BBKS transfer function [11] and Peacock-Dodds fitting formula for the nonlinear density power spectrum [11].

The optimal estimator.— Since $z \gtrsim 2$ galaxies are mainly lensed by matter at $z \lesssim 1$, $\kappa = A - B/\chi(z)$ is an excellent approximation, where $A$ and $B$ are two constants and $\chi$ is the comoving angular diameter distance. Since $\chi(z)$ varies slowly at $z > 2$, one can approximate $\kappa(\chi, \tilde{\eta}) \simeq (\kappa(\chi), \tilde{n})$, where $(\chi) = \sum \tilde{N}_i, \sum_j \tilde{N}_j$, $\tilde{N}_i$ is the effective distance to lens $i$. In the limit that $\tilde{N}_i \gg 1$, Poisson fluctuations become Gaussian. The likelihood function of $\kappa$ at an angular pixel, marginalized over $p(\delta_{g,11}, \cdots, \delta_{g,ij})$, the probability distribution of $\delta_{g,ij}$ of this angular pixel is,

$$L \propto \int \exp \left[ -\sum_{ij} \frac{[N_{ij} - \tilde{N}_{ij} - \tilde{N}_{ij}(W_{ij}\kappa + \delta_{g,ij})]^2}{2\tilde{N}_{ij}} \right] \times p(\delta_{g,11}, \cdots, \delta_{g,ij}) \prod_{ij} d\delta_{g,ij}.$$ 

(2)

We choose the redshift bin size $\Delta z \gtrsim 0.2$ and angular pixel size $\sim 10'$ such that $\delta_{g,ij}$ of different redshift bins are uncorrelated. For this choice, the matter density dispersion of each redshift bin $\sigma_m \lesssim 0.1$. This verifies the neglect of high order term $\delta_{g,k}$ in Eq. 1 & 2.

Since $\sigma_m \lesssim 0.1$ and galaxy bias $b_g$ is unlikely bigger than several [12], it is reasonable to assume that galaxies are Gaussian distributed. Then $p(\delta_{g,11}, \cdots)$ is completely determined by the covariance matrix $C_{ij} \equiv \langle \delta_{g,ij}(\tilde{n})\delta_{g,ij}(\tilde{n}) \rangle$. SKA can directly and accurately measure the correlations of galaxy density fluctuations between flux bins, which are the sum of $C_{ij}$, $C_{ij}$, correlations induced by lensing and cross terms. In the interesting range, $C_{ij}$ dominates. So one can take the measured sum as first guess of $C_{ij}$. Maximizing $L$, one obtains the optimal estimator $\tilde{\kappa}$ of $\kappa$. The reconstructed $\kappa$ can in turn be applied to subtract the lensing contribution in the covariance matrix estimation. This can be done iteratively. Since the lensing contribution is small, such iteration should be stable and converge quickly.

The properties of high redshift 21cm emitting galaxies are currently poorly known. It is likely that they trace the underlying dark matter at some level, and that galaxies of different luminosities are correlated to each other. We consider this case first, and then the extreme stochastic biasing limit [13] where galaxies of different flux are uncorrelated with each other. These two cases correspond to the worst and best cases for the $\kappa$ reconstruction, respectively.

Deterministic biasing.— We first consider the case that $\delta_{g,ij}$ of different flux bins (but of the same redshift bin) are linearly correlated, namely, $\delta_{g,ij} = b_g \delta_j$, where $\delta_j$ is the dark matter density of the $j$-th redshift bin. As discussed above, $b_{ij}$ can be measured iteratively. Marginalizing over $\delta_j$, we obtain

$$L \propto \exp \left[ -\frac{(\kappa - \tilde{\kappa})^2}{2(\Delta \kappa)^2} \right],$$

$$\tilde{\kappa} = \left( \sum_j S_j - \frac{B_j Q_j}{A_j} \right) (\Delta \kappa)^2,$$

$$\Delta \kappa = \left( \sum_j T_j - \frac{Q_j^2}{A_j} \right)^{-1/2},$$

(3)

Here $A_j = \sum_i (\tilde{N}_{ij} b_{ij})^2 / \tilde{N}_{ij} + 1/A_j^2$, $B_j = \sum_i (N_{ij} - \tilde{N}_{ij}) \tilde{N}_{ij} b_{ij} / \tilde{N}_{ij}$, $Q_j = \sum_i \tilde{N}_{ij} W_{ij} b_{ij} / \tilde{N}_{ij}$, $S_j = \sum_i (N_{ij} - \tilde{N}_{ij}) \tilde{N}_{ij} W_{ij} / \tilde{N}_{ij}$, and $T_j = \sum_i \tilde{N}_{ij} W_{ij}^2 / \tilde{N}_{ij}$. $\tilde{N}$ is the mean number of galaxies in each angular pixel. $\langle \cdots \rangle$ are weighted by galaxies with the noise from false peaks taken into account.

Maximal stochasticity.— Stochasticity eases the subtraction of the intrinsic clustering signal. In this case, $\delta_g$ of different bins are uncorrelated. We have

$$L \propto \exp \left[ -\sum_{ij} \frac{[N_{ij} - \tilde{N}_{ij} - \tilde{N}_{ij} W_{ij}\kappa]^2}{2\sigma_{ij}^2} \right],$$

$$\tilde{\kappa} = \frac{\sum_i (N_{ij} - \tilde{N}_{ij}) \tilde{N}_{ij} W_{ij} / \sigma_{ij}^2}{\sum_i \tilde{N}_{ij}^2 W_{ij}^2 / \sigma_{ij}^2},$$

$$\Delta \kappa = \left[ \partial^2 \ln L / \partial \kappa^2 \right]^{-1/2} = \left[ \sum_{ij} \frac{[\tilde{N}_{ij} W_{ij}]^2}{\sigma_{ij}^2} \right]^{-1/2},$$

$$\geq [\tilde{N} (W^2)]^{-1/2} \quad \text{when } \tilde{N}_{ij} \sigma_{ij}^2 \to 0.$$  

(4)

Here, $\sigma_{ij}^2 = \tilde{N}_{ij} + \tilde{N}_{ij}^2 \sigma_{ij}^2$, where the first term is the
non-Gaussian galaxy distribution should be taken into account. In two estimators for non-Gaussian galaxy distribution should be taken into account. Optimal estimators described above are no longer optimal. Optimal selection is conservative. We have neglected all systematics in this paper. (2) The galaxy bias. For the case of deterministic biasing, if \( b_g \propto W \), the signal disappears. This effect can be straightforwardly estimated through the \( \langle W^2 \rangle \) and \( \langle W \rangle \) terms in Eq. 3 & 4, once the HI mass function is measured. Since HI mass function at high \( z \) is effectively unknown, we postpone the discussion in this paper. (3) The galaxy distribution. When \( b_g \) is bigger than several or smoothing size is smaller than several arc-minutes, \( \delta_g \) is non-Gaussian. In this case, the estimators described above are no longer optimal. Optimal estimators for non-Gaussian galaxy distribution should be further investigated.

Applications.—The reconstructed \( \kappa \) map can be applied to many lensing statistics. For this purpose, reconstructed \( \kappa \) can be noisy because these statistics generally average over many angular pixels and achieve high \( S/N \). Then the optimal estimator derived in this paper can be applied to each narrow redshift bins and achieve high \( S/N \). Since \( n_g \) above \( 0.5-\sigma \) decreases much more slowly than \( t_{\text{int}} \) (for example, for the evolution model, decreasing \( t_{\text{int}} \) from 180 days/deg\(^2\) to 4 hours/deg\(^2\), \( n_g \) only decreases by a factor of 9), it is still likely to achieve \( S/N > 1 \) at \( \theta \sim 10' \) and scan rate of \( \sim 4000 \text{ deg}^2 \) per year (fig. 2). This will produce more lensing information \( (\times S/N \times f_{\text{sky}}^2) \) in a one year SKA survey than SNAP [23] will produce, which will cover 1000 deg\(^2\) sky area with \( \text{S/N} \sim 2 \) at smoothing scale \( \theta \sim 10' \).

Since the SKA will have \( \sim 0.3'' \) resolution at \( z \sim 2 \), it can resolve galaxies and measure cosmic shear. An intrinsic advantage of cosmic magnification measurement over cosmic shear measurement is that it does not require galaxies to be resolved. Thus, dwarf galaxies which are too small and too faint for reliable shear measurement still contribute to magnification measurement. Cosmic magnification exceeds cosmic shear at integration rate \( \lesssim 0.2-10 \text{ days/deg}^2 \) (fig. 1). We note that this comparison is conservative. We have neglected all systematics of shear measurement. For magnification estimation, we only select galaxies above a 0.5-\( \sigma \) detection threshold, or HI mass above several \( \times 10^8 M_\odot h^{-2} \). There are numerous galaxies with HI mass \( \sim 10^7 M_\odot h^{-2} \) [8], which can in principle be used to improve the measurement. We do not explore its potential in this paper since the luminosity function at the faint end is unclear.

Several uncertainties could degrade the signal separation. (1) The HI mass function, which is the dominant factor, as can be seen from table I and fig. 1. Here we further draw the attention on the slope of the HI mass function. For an extreme case that \( \alpha \rightarrow 1 \) and \( W \rightarrow 0 \) over a large flux range, the signal disappears. This effect can be straightforwardly estimated through the \( \langle W^2 \rangle \) and \( \langle W \rangle \) terms in Eq. 3 & 4, once the HI mass function is measured. Since HI mass function at high \( z \) is effectively unknown, we postpone the discussion in this paper. (2) The galaxy bias. For the case of deterministic biasing, if \( b_g \propto W \), flux information is no longer useful for the separation and our method effectively fails. But since \( b_g > 0 \), as long as the survey is deep enough to probe the faint end of galaxies where \( W < 0 \), \( b_g \) can not always mimic \( W \) and the separation is always possible. (3) The galaxy distribution. When \( b_g \) is bigger than several or smoothing size is smaller than several arc-minutes, \( \delta_g \) is non-Gaussian. In this case, the estimators described above are no longer optimal. Optimal estimators for non-Gaussian galaxy distribution should be further investigated.

Results.—SKA is able to detect \( n_g \gtrsim 100 \text{ arcmin}^{-2} \) at \( z \gtrsim 2 \) (table I). For an integration time \( t_{\text{int}} = 18 \text{ days/deg}^2 \), a \( S/N \gtrsim 2 \) can be achieved at \( \theta \sim 10' \) (fig. 1). Deep survey configuration detects more faint galaxies, which have \( W \rightarrow -2 \), mimic a constant \( b \) and thus do not contribute to the signal, due to the \( \langle W^2 \rangle - \langle W \rangle^2/(b^2) \) factor in Eq. 3. An optimal survey configuration should have \( \langle W \rangle \rightarrow 0 \), which can be achieved at \( t_{\text{int}} \sim 0.2-1 \text{ day/deg}^2 \) (fig. 1). Since \( n_g \) above \( 0.5-\sigma \) decreases much more slowly than \( t_{\text{int}} \) (for example, for the evolution model, decreasing \( t_{\text{int}} \) from 180 days/deg\(^2\) to 4 hours/deg\(^2\), \( n_g \) only decreases by a factor of 9), it is still likely to achieve \( S/N > 1 \) at \( \theta \sim 10' \) and scan rate of \( \sim 4000 \text{ deg}^2 \) per year (fig. 2). This will produce more lensing information \( (\times S/N \times f_{\text{sky}}^2) \) in a one year SKA survey than SNAP [23] will produce, which will cover 1000 deg\(^2\) sky area with \( \text{S/N} \sim 2 \) at smoothing scale \( \theta \sim 10' \).

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allows the lensing tomography. (1) The probability density function $p(\kappa)$. $p(\kappa)$ as a function of $\kappa$ and smoothing angular size $\theta$ can provide independent constraints on cosmology. Recently [15] showed that the Wiener filter reconstruction of $p(\kappa)$ from noisy convergence map can do deep into regions where $|\kappa|/\Delta\kappa \ll 1$. We thus expect that $p(\kappa)$ can be recovered accurately from SKA.

(2) Lensing power spectrum and bispectrum. The reconstructed $\kappa$ map barely has $S/N \gtrsim 5$, so it is consistent to neglect $\lesssim 10\%$ higher order terms: $O(\kappa^2)$ terms and $\delta \kappa$ term neglected in Eq. 1 and $\kappa(\chi) - \langle \kappa \rangle$. But these terms should be taken into account for precision measurement of lensing power spectrum and bispectrum, since their statistical errors can reach $\sim 1\%$ accuracy [16]. For the linear estimator we derived, contributions of these terms to the power spectrum and bispectrum can be straightforwardly and robustly predicted. So, there is no need to derive a more complicated nonlinear estimator. (3) Cluster finding and cluster density profile. This is a promising approach to break the cluster mass sheet degeneracy. In the reconstructed maps, massive clusters at $z \sim 0.2$ show as high peaks with strength $\kappa \sim 0.1$ and size $\sim 10^5$ and can be easily identified. These clusters are excellent objects to measure the geometry of the universe by the technique of lensing cross-correlation tomography [17]. Since $S/N$ is so high, one can choose smoothing size $\sim 1$ and measure the projected cluster density profile. Exerting a prior on cluster density profile, the reconstruction can be further improved [17]. When $\kappa \rightarrow 1$, the weak lensing condition breaks and Eq. 4 no longer holds. By utilizing the exact magnification equation, one can develop new estimator, in analogy to the reduced shear reconstruction [18]. We leave this topic for further study.

We summarize our results. Cosmic magnification is statistically more sensitive than cosmic shear because it is possible to use the large number of galaxies detected at low statistical significance. Intrinsic clustering can be subtracted because (1) magnification depends strongly on the shape of the luminosity function, which varies significantly, while intrinsic clustering depends weakly on the intrinsic luminosity itself and (2) they have different redshift dependence. Cosmic magnification shows promise as a complementary technique to map the statistically precise distribution of matter, which is not subject to most of the systematics of cosmic shear. We have worked through the specific numbers for the SKA, but the general formalism would also apply to optical spectroscopic or photometric redshift surveys.

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[1] A. Refregier, 2003, ARAA, 41, 645 and reference therein; M. Jarvis et al. 2003, AJ, 125, 1014; U. Pen et al. 2003, 592, 664; M. Jarvis et al. 2004, MNRAS, 352, 338; T. Chang et al. 2004, astro-ph/0408548
[2] C. Harata et al. 2004, MNRAS, 353, 529; H. Hoekstra, 2004, MNRAS, 347, 1337; M. Jarvis et al. 2004, MNRAS, 352, 338; C. Vale et al. 2004, ApJL, 613, L1; L. van Waerbeke et al. 2004, astro-ph/0406468; C. Heymans et al. 2005, astro-ph/0506112
[3] U. Seljak and M. Zaldarriaga, 1999, PRL, 82, 2636; M. Zaldarriaga and U. Seljak, 1999, PRD, 5913507; W. Hu and T. Okamoto, 2002, ApJ, 574, 566; C. Hirata and U. Seljak, 2003, PRD, 68, 083002
[4] U. Pen, 2004, New Astronomy, 9, 417; K. Sigurdson and A. Cooray, 2005, astro-ph/0502549
[5] A. Amblard et al. 2004, New Astronomy, 9, 687
[6] R. Scranton et al., 2005, astro-ph/0504510 and reference therein
[7] P. Zhang and U. Pen, 2005, astro-ph/0504551
[8] P. Schneider et al. 1992, Gravitational lenses, Springer-Verlag, Berlin
[9] M. Zwaan et al. 1997, ApJ, 490, 173
[10] J. Bardeen et al. 1986, ApJ, 304, 15
[11] J. Peacock and S. Dodds, 1997, MNRAS, 280, 19
[12] H. Mo and S. White, 1996, MNRAS, 282, 347; M. Gavvalisco et al. 1998, ApJ, 503, 543; V. Springel et al., 2005, Nature, 435, 629
[13] U. Pen, 1998, ApJ, 504, 601.
[14] B. Ménard and M. Bartelmann, 2002, A&A, 386, 784
[15] T. Zhang and U. Pen, 2005, astro-ph/0503064
[16] B. Jain and A. Taylor, 2003, PRL, 91, 141302; J. Zhang et al. 2003, astro-ph/0312348
[17] S. Dodelson, 2003, PRD, 70, 023009
[18] U. Pen, 2000, ApJ, 534, L19
[19] The observed $dn/dF$ is convolved with system noise. Because there are more dwarf galaxies than massive ones, noise makes the observed $dn/dF$ both larger and steeper, in the flux range that SKA can probe at $z \gtrsim 2$. The overall effect is that system noise in flux measurements increases the cosmic magnification signal and strengthens the result in this paper. For simplicity, we neglect this complexity.
[20] Cosmic magnification does not change the averaged galaxy spatial and flux distribution, up to $O(\kappa^2) \sim 10^{-4}$ accuracy. The sky coverage of SKA is $\gtrsim 100$ deg$^2$. Thus for each redshift and flux bin, there are $\gtrsim 4000$ angular pixels with size $\theta \sim 10^5$ and $\gtrsim 10^5$ galaxies across the survey sky, so $N_i$ can be measured accurately. $\tilde{N}_ij$ can be accurately predicted, since system noise is Gaussian and the dispersion $S_{\text{sys}}$ is specified for each survey. The number of false peaks with flux above $n-\sigma$, or $nS_{\text{sys}}$ per redshift interval per beam is

$$[1.4 \, \text{GHz}/\Delta v(1+z)^2] \text{Erf}[n/\sqrt{2}]/2. \Delta v \text{ is chosen to be the frequency width corresponding to 100 km/s velocity dispersion at redshift } z.$$ Thus, one can accurately predict $N^S_{ij}$ and $W_{ij}$.

[21] SKA http://www.skatelescope.org/
[22] The approximation $\kappa \simeq \langle \kappa \rangle$ simplifies the derivation of the optimal estimator significantly, though its accuracy can be as bad as $\sim \pm 20\%$, at each redshift bins. But after
averaging over many redshift bins, corrections in different bins effectively cancel. For the optimal estimators derived (Eq. 3 & 4), one can derive the unbiased expression of \( \langle \kappa \rangle \) such that \( \langle \kappa \rangle = \langle \hat{\kappa} \rangle \).

[23] SNAP: http://snap.lbl.gov/