Clinical efficacy of icotinib in lung cancer patients with different EGFR mutation status: a meta-analysis

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ABSTRACT

Icotinib is a novel and the third listed epidermal growth factor receptor-tyrosine kinase inhibitors (EGFR-TKIs), which exerts a good anti-tumor efficacy on non-small cell lung cancer (NSCLC). The efficacy of EGFR-TKIs has been shown to be associated with the EGFR mutation status, especially exon 19 deletion (19Del) and exon 21 L858R mutation. Therefore, a meta-analysis was performed to assess the efficacy of icotinib in NSCLC patients harboring EGFR mutations (19Del or L858R) and wild type (19Del and L858R loci wild type). A total of 24 studies were included for comparing the objective response rate (ORR) in the EGFR wild type and mutant patients treated with icotinib. The ORRs of EGFR mutant patients (19Del or L858R) are better than those of EGFR wild type patients (OR = 7.03(5.09–9.71), \(P < 0.00001\)). The pooling ORs from 21 studies on the disease control rate (DCR) in EGFR mutant patients are better than those of EGFR wild type patients (OR = 10.54(5.72–19.43), \(P < 0.00001\)). Moreover, the ORRs of EGFR 19Del patients are better than those of EGFR L858R patients after pooling ORs of 12 studies (OR = 2.04(1.12–3.73), \(P = 0.019\)). However, there was no significant difference on DCRs of EGFR 19Del patients and those of EGFR L858R patients (OR = 2.01(0.94–4.32), \(P = 0.072\)). Our findings indicated that compared with EGFR wild type patients, EGFR mutant patients have better ORRs and DCRs after icotinib treatment; EGFR 19Del patients treated with icotinib have better ORRs than EGFR L858R patients. EGFR mutation status is a useful biomarker for the evaluation of icotinib efficacy in NSCLC patients.

INTRODUCTION

Lung cancer is the leading cause of mortality around the world and non-small cell lung cancer (NSCLC) is up to 85% of all types of lung cancer [1]. NSCLC mainly includes squamous cell carcinoma, adenocarcinoma and large cell carcinoma. No obvious clinical manifestations were observed at early stage and more than 40% of NSCLC are metastatic (Stage IV) disease at diagnosis [2].

With in-depth studies of genesis and cancer related signal pathway, epidermal growth factor receptor (EGFR)-dependent pathway was revealed to play important roles in the development and progression of epithelial cells in NSCLC patients [3]. EGFR tyrosine kinase inhibitors (EGFR-TKIs) play important roles in the treatment of advanced NSCLC because of their superior efficacy over than chemotherapy [4]. EGFR-TKIs such as gefitinib and erlotinib were identified to extend survival and increase quality of life in NSCLC patients [4–6]. Icotinib was a novel and the third listed EGFR-TKIs, which could exert a good anti-tumor efficacy on NSCLC [7], especially in the re-treatment of advanced NSCLC [8]. Up to now, it
has become one of the standard drugs for the treatment of advanced NSCLC in China [9].

The efficacy of EGFR-TKIs has been shown to be associated with the EGFR mutation status, especially exon 19 and exon 21 that are sensitive to targeted drug therapies [4, 7, 10–15]. Moreover, previous studies has revealed that patients treated with icotinib harboring EGFR exon 19 deletion (19Del) had better survival than those harboring exon 21 point mutation (L858R) [16] or there were no difference among the patients harboring 19Del or L858R mutations [7, 17, 18]. Because of the inconsistent results, relative small sample sizes and lack of high quality studies, their conclusions are limited value. Therefore, we reviewed all the publications about icotinib and conducted a meta-analysis to assess the efficacy of icotinib in NSCLC patients harboring EGFR mutations (19Del or L858R) or wild type for these two mutations.

RESULTS

Study review and selection

The study selection procedure is shown in Figure 1. We searched from the databases including PubMed, EMBASE, Web of Science, Wanfang and Chinese National Knowledge Infrastructure (CNKI) to 14th Oct. 2016. A total of 136 publications were found after excluding the duplicated studies. Then we excluded 62 irrelevant studies, 5 meta-analysis, 9 case reports and 12 basic studies. Forty-eight studies were included for further review. We further excluded the article that has no EGFR status or no clinical indicators including objective response rate (ORR) and disease control rate (DCR). Twenty-four publications having EGFR mutation and wide type data were selected to qualitative synthesis. Fifteen studies having EGFR 19Del and L858R mutation data were included in qualitative synthesis. Finally, 24 studies having ORR values and 21 studies having DCR values in EGFR mutant and wild type patients were enrolled for meta-analysis. Twelve studies having ORR values and 8 studies having DCR values in EGFR 19Del and L858R patients were enrolled for meta-analysis.

Characteristics of included studies

Not all studies had all clinical indicators data. After summarized, there were 24 publications have ORR data compared with EGFR mutation and wild type patients. Among the 24 publications, except one publication [19], 23 publications have DCR data. There were 12 publications having ORR values and 8 publications having DCR values in EGFR 19Del and L858R patients (Table 1). The characteristics of first author’s name, publishing year, region, Newcastle-Ottawa Scale (NOS) score, study design and the number of patients harboring EGFR mutation status were shown in Table 1.

Quality evaluation of enrolled publications

The details of the quality evaluation of enrolled publications were shown in Figure 2. There were no randomized controlled trials involved in EGFR status on the efficacy of icotinib in NSCLC patients harboring different EGFR mutation status. All studies were cohort trials and retrospective studies. Therefore the risk of bias was high regarding adequate sequence generation and blinding. But other methodological issues present relatively little risk.

EGFR wild type vs. EGFR mutation

There were 24 publications enrolled for comparing the ORR in EGFR wild type patients and EGFR mutant patients (19Del or L858R). Total patients enrolled in the meta-analysis were 1300 including 825 EGFR mutant patients and 475 EGFR wild type patients. There was no heterogeneity among 24 publications (I² = 5.4%, P = 0.387). Therefore, Mantel-Haenszel fixed effects model was used to calculate the pooled odds ratio of included studies. The results showed that the ORRs of EGFR mutant patients are better than those of EGFR wild type patients (OR = 7.03 (5.09–9.71), P < 0.00001) (Figure 3A).

Among the 24 publications, one publication [19] has no data for DCRs; DCRs were 100% in EGFR mutant patients and EGFR wild type patients in two publications [12, 20]. After excluding these 3 publications, 21 publications having 1127 patients including 750 EGFR mutant patients and 377 EGFR wild type patients were enrolled in meta-analysis for DCRs. Because of heterogeneity (I² = 53.2%, P = 0.002), Mantel-Haenszel random effects model was used to analyze. The results showed that the DCRs of EGFR mutant patients are better than those of EGFR wild type patients (OR = 10.54 (5.72–19.43), P < 0.00001) (Figure 3B). According to subgroup analysis for NOS, there was no heterogeneity in NOS 5 and 6 groups (I² = 0.0%, P = 0.587; I² = 0.0%, P = 0.574, respectively), but NOS 7 groups has heterogeneity (I² = 62.4%, P = 0.021). The DCRs of EGFR mutant patients are better than those of EGFR wild type patients according to subgroups for different NOS (OR = 5.06(1.37–18.7), P = 0.015; 14.53(8.47–24.9), P < 0.0001; 10.37(2.25–42.56), P = 0.001, respectively) (Figure 3C).

EGFR 19Del vs. EGFR L858R

There were 12 publications enrolled for comparing the ORR in EGFR 19Del and EGFR L858R patients. Total patients enrolled in the meta-analysis on ORRs were 555, including 303 EGFR 19Del patients and 252 L858R patients. Because of heterogeneity (I² = 49.7%, P = 0.025), a Mantel-Haenszel random-effects model was used to analyze. The results showed that the ORRs...
of EGFR 19Del patients are better than those of EGFR L858R patients (OR = 2.04 (1.12–3.73), P = 0.019) (Figure 4A). After excluded one publication [21] that could influence the overall effective size, there was no heterogeneity (I² = 43.8%, P = 0.058) and the ORRs of EGFR 19Del patients are better than those of EGFR L858R patients (OR = 1.48 (1.02–2.13), P = 0.037) (Figure 4B).

Among the 12 publications, one publication [7] has no data for DCRs; DCRs were 100% in EGFR 19Del and L858R patients in three publications [12, 20, 22]. After excluded these 4 publications, 8 publications containing 392 patients including 210 EGFR 19Del patients and 182 L858R patients were enrolled in meta-analysis for DCRs. Because there was no heterogeneity (I² = 0.0%, P = 0.776), a Mantel-Haenszel fixed-effects model was used to analyze. The results showed that there were no significant differences on DCRs of EGFR 19Del patients and L858R patients (OR = 2.01 (0.94–4.32), P = 0.072) (Figure 4C).

Publication bias and sensitivity analysis

Publication bias was examined by funnel plots, Egger’s test and Begg’s test. For pooling ORs analysis for ORRs in EGFR wild type patients and mutant patients, funnel plots, Egger’s test (P = 0.09) and Begg’s test (P = 0.747) showed no publication bias (Figure 5A–5C respectively). For pooling ORs analysis for DCRs in EGFR wild type patients and mutant patients, funnel plots

Table 1: Characteristics of studies included in meta-analysis

| Author          | Year | Region | Study design | NOS Score | EGFR mutant and wild type | EGFR 19Del and L858R | References |
|-----------------|------|--------|--------------|-----------|---------------------------|----------------------|------------|
| Zhao Q          | 2011 | China  | R            | 6         | M/W number               | 5/14                 | Y Y        | [20]       |
| Zen Xiao-Mei    | 2013 | China  | R            | 6         | 15/3                     | Y Y                  | N N        | [38]       |
| Yang Xin-Jie    | 2013 | China  | R            | 6         | 18/2                     | Y Y                  | 10/8       | [39]       |
| Yang S          | 2016 | China  | R            | 7         | 32/55                    | Y N                  | N N        | [31]       |
| Wei Feng-Lu     | 2015 | China  | R            | 6         | 15/15                    | Y Y                  | N N        | [40]       |
| Wang Tao        | 2016 | China  | R            | 7         | 38/29                    | Y Y                  | 15/16      | [32]       |
| Tan Fen-Lai     | 2012 | China  | R            | 7         | 249/101                  | Y Y                  | 72/88      | [35]       |
| Sun Jing        | 2014 | China  | R            | 7         | 34/10                    | Y Y                  | N N        | [41]       |
| Song Zheng-Bo   | 2013 | China  | R            | 6         | 36/13                    | Y Y                  | N N        | [42]       |
| Shao L          | 2014 | China  | R            | 7         | 12/16                    | Y Y                  | N N        | [43]       |
| Ren G-J         | 2011 | China  | R            | 5         | 7/7                      | Y Y                  | 3/4        | [44]       |
| Pang Lin-Rong    | 2014 | China  | R            | 6         | 33/18                    | Y Y                  | 19/13      | [21]       |
| Nong Jin-Yin    | 2013 | China  | R            | 6         | 23/9                     | Y Y                  | 14/9       | [22]       |
| Na Qin          | 2013 | China  | R            | 6         | 35/11                    | Y Y                  | N N        | [16]       |
| Ma Xiang-Leiei  | 2014 | China  | R            | 7         | 40/14                    | Y Y                  | N N        | [45]       |
| Li Xing-Ping-Ping| 2015 | China  | R            | 5         | 10/2                     | Y Y                  | N N        | [46]       |
| Li Xi           | 2012 | China  | R            | 6         | 23/36                    | Y Y                  | 55/40      | [34]       |
| Li Xi           | 2015 | China  | R            | 6         | 99/25                    | Y Y                  | 13/9       | [18]       |
| Li Ran          | 2013 | China  | R            | 5         | 23/7                     | Y Y                  | 17/6       | [47]       |
| He Xiao-Tin     | 2015 | China  | R            | 6         | 33/13                    | Y Y                  | N N        | [48]       |
| He Chun-Xiao    | 2012 | China  | R            | 7         | 1/4                      | Y Y                  | N N        | [49]       |
| Guo Lei         | 2016 | China  | R            | 7         | 21/6                     | Y Y                  | N N        | [50]       |
| Gu A            | 2013 | China  | R            | 5         | 4/2                      | Y Y                  | N N        | [51]       |
| Chen Xiao-Feng  | 2014 | China  | R            | 6         | 19/63                    | Y Y                  | N N        | [33]       |
| Zhang Xiao-Xue  | 2016 | China  | R            | 7         | N N                      | 60/44                | Y N        | [7]        |
| Shen Yan-Wei    | 2016 | China  | R            | 6         | N N                      | 21/14                | Y Y        | [17]       |

Abbreviations: Y, Yes; N, NO; M, Mutation; W, wild type; Del, deletion; R, Retrospective study; NOS, Newcastle-Ottawa Scale; ORR, Objective Response Rate; DCR, disease control rate.
and Begg’s ($P = 0.651$) test showed no publication bias (Figure 5D and 5F, respectively). However, Egger’s test ($P = 0.003$) showed publication bias (Figure 5E).

For pooling ORs analysis for ORRs among $EGFR$ 19Del and L858R patients, funnel plots, Egger’s test ($P = 0.164$) and Begg’s test ($P = 0.451$) showed no publication bias (Figure 6A–6C, respectively). For pooling ORs analysis for DCRs in the $EGFR$ 19Del and L858R patients, funnel plots, Egger’s test ($P = 0.376$) and Begg’s test ($P = 0.711$) showed no publication bias (Figure 6D–6F, respectively).

Sensitivity analysis results showed that changing the effect models had no significant effects on the pooled ORs and the final strength of the association between $EGFR$ polymorphisms and clinical efficacy of icotinib in lung cancer patients. Moreover, Figure 7A showed the results of sensitivity analysis regarding DCRs of $EGFR$ mutant patients vs. $EGFR$ wild type patients. We found that excluded studies did not influence the overall effective size in DCRs analysis of $EGFR$ mutant patients vs. $EGFR$ wild type patients. Figure 7B showed one publication (Pang Lin-Rong, 2014) [21] could influence the overall effective size in ORRs analysis of $EGFR$ 19Del and $EGFR$ L858R patients. After we excluded this publication [21], There was no heterogeneity for pooling ORs analysis for ORRs in $EGFR$ 19Del and $EGFR$ L858R patients ($I^2 = 43.8\%$, $P = 0.058$) (Figure 3B).

**DISCUSSION**

We carried out this meta-analysis to evaluate the clinical efficacy of icotinib in lung cancer patients with different $EGFR$ mutation status. Our meta-analysis suggests that compared with $EGFR$ wild type patients, $EGFR$ mutant patients have better ORRs and DCRs after icotinib treatment; compared with $EGFR$ L858R patients, $EGFR$ 19Del patients have better ORRs after icotinib treatment.

It is widely accepted that $EGFR$ plays important roles in tumor cell proliferation, invasion, angiogenesis, metastasis and apoptosis [23]. $EGFR$-TKIs such as gefitinib or erlotinib inhibit tumor cells by blocking the $EGFR$ signaling via binding to ATP binding site of $EGFR$ to increase survival in NSCLC patients [8, 24, 25]. Several clinical studies have confirmed that patients with $EGFR$ mutation have benefited from $EGFR$-TKIs treatment [26–28]. There were two major $EGFR$-activating
Random sequence generation (selection bias)
Allocation concealment (selection bias)
Blinding of participants and personnel (performance bias)
Blinding of outcome assessment (detection bias)
Incomplete outcome data (attrition bias)
Selective reporting (reporting bias)
Other bias

Figure 2: Analysis of risk of bias.

**A**

| Study ID | OR (95% CI) | % Weight |
|----------|-------------|----------|
| Zhou G (2011) | 1.50 (0.90, 2.50) | 4.05 |
| Zue X-M (2013) | 3.00 (3.10, 10.30) | 1.97 |
| Yang J-L (2013) | 5.32 (5.09, 5.50) | 0.87 |
| Yang S (2016) | 3.09 (1.24, 7.98) | 16.89 |
| Wei P-L (2015) | 13.00 (2.57, 61.98) | 1.97 |
| Tan (2012) | 0.02 (0.02, 4.73) | 3.00 |
| Sun (2012) | 4.47 (1.27, 1.51) | 20.00 |
| Song Z-B (2012) | 1.73 (1.35, 1.30) | 1.19 |
| Shao J (2012) | 18.80 (1.07, 143.49) | 1.19 |
| Ren J (2011) | 11.67 (4.88, 28.02) | 0.63 |
| Pang L-R (2010) | 18.40 (3.15, 95.92) | 2.32 |
| Ning J-Y (2010) | 13.50 (1.68, 141.17) | 1.46 |
| Ni D (2010) | 13.33 (2.03, 118.50) | 1.60 |
| Ma K-L (2014) | 14.03 (2.71, 72.99) | 2.63 |
| Liang W-F (2011) | 0.43 (0.20, 9.96) | 5.46 |
| Li X (2013) | 7.37 (1.49, 34.08) | 6.01 |
| Li R (2013) | 4.00 (0.40, 40.61) | 1.72 |
| He X-T (2013) | 2.00 (0.30, 13.50) | 0.21 |
| He N-Z (2012) | 14.09 (0.71, 29.74) | 1.07 |
| Gu J (2015) | 5.00 (1.10, 23.00) | 0.92 |
| Gu A (2013) | 13.71 (4.05, 48.41) | 3.53 |
| Chen K (2014) | 7.03 (3.50, 17.71) | 100.00 |

**B**

| Study ID | OR (95% CI) | % Weight |
|----------|-------------|----------|
| Zuo X-M (2013) | 7.00 (3.30, 160.20) | 2.84 |
| Yang X-J (2013) | 17.00 (1.70, 252.79) | 2.48 |
| Wei P-L (2012) | 11.00 (4.08, 205.07) | 6.01 |
| Tan F-L (2012) | 4.00 (1.24, 7.98) | 16.89 |
| Sun J (2014) | 18.80 (1.07, 143.49) | 1.19 |
| Song Z-B (2012) | 1.73 (1.35, 1.30) | 1.19 |
| Shao J (2012) | 18.80 (1.07, 143.49) | 1.19 |
| Ren J (2011) | 11.67 (4.88, 28.02) | 0.63 |
| Pang L-R (2010) | 18.40 (3.15, 95.92) | 2.32 |
| Ning J-Y (2010) | 13.50 (1.68, 141.17) | 1.46 |
| Ni D (2010) | 13.33 (2.03, 118.50) | 1.60 |
| Ma K-L (2014) | 14.03 (2.71, 72.99) | 2.63 |
| Liang W-F (2011) | 0.43 (0.20, 9.96) | 5.46 |
| Li X (2013) | 7.37 (1.49, 34.08) | 6.01 |
| Li R (2013) | 4.00 (0.40, 40.61) | 1.72 |
| He X-T (2013) | 2.00 (0.30, 13.50) | 0.21 |
| He N-Z (2012) | 14.09 (0.71, 29.74) | 1.07 |
| Gu J (2015) | 5.00 (1.10, 23.00) | 0.92 |
| Gu A (2013) | 13.71 (4.05, 48.41) | 3.53 |
| Chen K (2014) | 7.03 (3.50, 17.71) | 100.00 |

**C**

| Study ID | OR (95% CI) | % Weight |
|----------|-------------|----------|
| Zuo X-M (2013) | 7.00 (3.30, 160.20) | 2.84 |
| Yang X-J (2013) | 17.00 (1.70, 252.79) | 2.48 |
| Wei P-L (2012) | 11.00 (4.08, 205.07) | 6.01 |
| Tan F-L (2012) | 4.00 (1.24, 7.98) | 16.89 |
| Sun J (2014) | 18.80 (1.07, 143.49) | 1.19 |
| Song Z-B (2012) | 1.73 (1.35, 1.30) | 1.19 |
| Shao J (2012) | 18.80 (1.07, 143.49) | 1.19 |
| Ren J (2011) | 11.67 (4.88, 28.02) | 0.63 |
| Pang L-R (2010) | 18.40 (3.15, 95.92) | 2.32 |
| Ning J-Y (2010) | 13.50 (1.68, 141.17) | 1.46 |
| Ni D (2010) | 13.33 (2.03, 118.50) | 1.60 |
| Ma K-L (2014) | 14.03 (2.71, 72.99) | 2.63 |
| Liang W-F (2011) | 0.43 (0.20, 9.96) | 5.46 |
| Li X (2013) | 7.37 (1.49, 34.08) | 6.01 |
| Li R (2013) | 4.00 (0.40, 40.61) | 1.72 |
| He X-T (2013) | 2.00 (0.30, 13.50) | 0.21 |
| He N-Z (2012) | 14.09 (0.71, 29.74) | 1.07 |
| Gu J (2015) | 5.00 (1.10, 23.00) | 0.92 |
| Gu A (2013) | 13.71 (4.05, 48.41) | 3.53 |
| Chen K (2014) | 7.03 (3.50, 17.71) | 100.00 |

Figure 3: Forest plots of studies evaluating odds ratios of ORRs (A), DCRs (B) and DCRs subgroups analysis according to NOS (C) in EGFR wild type and EGFR mutant patients. OR: odds ratio.
**Figure 4:** Forest plots of studies evaluating odds ratios of ORRs (A), forest plots of studies excluding one publication that influences the overall effective size evaluating odds ratios of ORRs (B) and DCRs (C) in EGFR 19Del and L858R patients. OR: odds ratio.

**Figure 5:** The funnel plots, Egger’s test and Begg’s test of publication bias in pooling ORs analysis in EGFR wild type and EGFR mutant patients. (A) The funnel plots, (B) Egger’s test and (C) Begg’s test for pooling ORs of ORRs analysis; (D) The funnel plots, (E) Egger’s test and (F) Begg’s test for pooling ORs of DCRs analysis. OR: odds ratio; SE: standard error.
mutations including an in-frame deletion in exon 19 and an L858R substitution in exon 21, which account for about 90% of all clinically important mutations related to EGFR-TKIs sensitivity [29, 30]. Icotinib is a new type and the third world-listed EGFR-TKIs to treat for NSCLC patients [8]. Several studies have revealed that EGFR mutant patients could get more benefit from icotinib treatment than EGFR wild type patients [18, 22, 31–33]. More evidences have the trend to imply that EGFR 19Del patients have better efficacy of icotinib than L858R patients [20, 34, 35]. However, there were some negative results against these views [7, 14, 17]. Furthermore, these studies were small sample size and retrospective study. Whether EGFR mutation status, especial EGFR 19Del and L858R, influence the efficacy of icotinib in NSCLC patients, it is still unclear now. Therefore, we carried out this meta-analysis to figure them out.

Icotinib is a new EGFR-TKI used in the treatment of NSCLC patients and it is only available in China. There were no randomized controlled trials on the efficacy of icotinib in NSCLC patients harboring different EGFR mutation status. Most of studies were cohort and retrospective studies enrolled small sample sizes. Therefore, there were some risks of bias such as high regarding adequate sequence generation and blinding.

The heterogeneities of studies are common in meta-analysis because of the different type of studies, methods or the selected case and controls. In our study, there was no

Figure 6: The funnel plots, Egger’s test and Begg’s test of publication bias in pooling ORs analysis in EGFR 19Del and L858R patients. (A) The funnel plots, (B) Egger’s test and (C) Begg’s test for pooling ORs of ORRs analysis; (D) The funnel plots, (E) Egger’s test and (F) Begg’s test for pooling ORs of DCRs analysis. OR: odds ratio; SE: standard error.

Figure 7: The sensitivity analysis of pooling ORs of DCRs in EGFR mutant patients vs. EGFR wild type patients (A) and ORRs in EGFR 19Del and EGFR L858R patients (B). CI: confidence interval.
heterogeneity when we pooling the ORs of ORRs in EGFR wild type and mutant patients, the ORs of DCRs in EGFR 19Del and L858R patients. Meta-analysis results showed significant between-study heterogeneity in pooling analysis about DCRs in EGFR wild type and mutant patients and the ORRs in EGFR 19Del and L858R patients. Therefore, Mantel-Haenszel random-effects model was used for analysis. In order to figure out the source of heterogeneity in pooling ORs of DCRs in EGFR wild type and mutant patients, we conducted subgroup analysis by NOS. After the subgroup analysis by NOS, there was no heterogeneity in NOS 6 and 5 groups, but there was heterogeneity in NOS 7 group (Figure 3C). In order to draw more cautious conclusion on EGFR status on icotinib efficacy in NSCLC patients, we also performed the sensitivity analysis. We found that excluded studies did not influence the overall effective size in DCRs analysis of EGFR mutant patients vs. EGFR wild type patients (Figure 7A). For heterogeneity of the pooling ORRs in EGFR 19Del and L858R patients, sensitivity analysis showed one publication (Pang Lin-Rong, 2014) [21] could influence the overall effective size in ORRs analysis of EGFR 19Del and EGFR L858R patients(Figure 7B). After we excluded this publication [21], there was no heterogeneity for pooling ORs analysis for ORRs in EGFR 19Del and EGFR L858R patients (P = 43.8%, P = 0.058) (Figure 3B). We also carried the publication bias analysis by funnel plots, Egger’s test and Begg’s test. The results showed there was publication bias for pooling ORs analysis for DCRs in the EGFR wild type patients and EGFR mutant patients. Negative results of studies may not be published. Other three groups of pooling ORs analysis showed no publication bias. The inconsistency of these studies may be due to source of patients, disease condition, publication qualities or other clinical issues. Further large sample multi-center studies and well-designed research are needed.

There were some limitations in this meta-analysis. First, the studies enrolled in our meta-analysis were limited. Slight publication bias may exist because the research having negative results may have not been published online. Second, most studies included in our meta-analysis were retrospective and cohort studies, which may have selection bias and blinding bias. The relative small subjects in studies also influence the meta-analysis results.

In conclusion, our meta-analysis suggests that compared with EGFR wild type patients, EGFR mutant patients have better ORRs and DCRs after icotinib treatment; EGFR 19Del patients have better ORRs after icotinib treatment than EGFR L858R patients. EGFR mutation status is a useful biomarker for evaluation of icotinib efficacy in NSCLC patients. Because of limited numbers of studies and small sample sizes included as well as the heterogeneity in our meta-analysis, more randomized and large-scale clinical trials are necessary to confirm our meta-analysis results.

**MATERIALS AND METHODS**

**Study review and selection**

We reviewed the databases including PubMed, EMBASE, Web of Science, Wanfang and CNKI to 14 Oct. 2016. The searching strategy was “icotinib” or “conmana” and “cancer or carcinoma or tumor”. Dr. Jian Qu and Dr. Ya-Nan Wang reviewed all relevant articles to identify potential eligible studies.

**Inclusion and exclusion criteria**

The inclusion criteria were including as follows: clinical study on NSCLC patients harboring EGFR status (EGFR wild type is defined as no mutation for 19Del and L858R; EGFR mutation is defined as harboring EGFR 19Del or L858R mutation) using icotinib treatment; at least have the one clinical indicator (ORR and DCR). A study was excluded if it was not relevant to cancer and clinical patients, EGFR status, or had no clinical indicators; involved just in animals or cells; or was a review, or abstract having no data. Different opinions on selections were solved by all author’s discussion.

**Data collection, quality assessment and assessment of risk of bias.**

Therapeutic efficacy was evaluated by response evaluation criteria in solid tumors (RECISTC) including complete response (CR), partial response (PR), stable disease (SD) and progressive disease (PD). Clinical outcome indicators include objective response rate (ORR) and disease control rate (DCR). The data on authors’ names, sex, smoking status. Two investigators (Dr. Qu Jian and Dr. Ya-Nan Wang) independently extracted EGFR mutation type, numbers and clinical outcomes. The Newcastle-Ottawa Scale (NOS) method was used to evaluate the quality of selected studies. According to the Cochrane Handbook version 5.1.0 [36], the risk of bias was assessed including method of random sequence generation (selection bias), allocation concealment (selection bias), blinding (performance bias and detection bias), incomplete outcome data (detection bias), and selective reporting (detection bias). We evaluated methodological quality as low, high, or unclear risk of bias.

**Statistical analysis**

We used STATA version 12 (Stata Corp, College Station, TX, USA) to carry out the meta-analysis. Heterogeneity was assessed by the Cochrane’s Q-statistic test and F test. We used a Mantel-Haenszel random effects model in the analysis if $P < 0.05$ and $I^2 > 50\%$, otherwise, a Mantel-Haenszel fixed effects model was chosen [37]. Count data calculate the relative risk (RR)
or odds ratio (OR) and expressed as 95% confidence interval (CI). Publication bias was analyzed by funnel plot, Egger’s test and Begg’s test. Sensitivity analysis was performed to assess the influence of a single study on the overall effective size. Tests were two-sided and statistical significance was accepted at $P < 0.05$.

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**CONFLICTS OF INTEREST**

The authors declare no conflicts of interest.

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