Direct vs. Indirect Impact of Fuel Price on Wages in South Africa: An Analysis Using the Impulse-Response Function

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Abstract

The international oil price has been increasing significantly, but consumers react more to the announcement of the increase in fuel (pump) price, than an increase in international oil prices. This study, therefore, investigates the direct effect of fuel prices on the inflation rate and wages in South Africa, on the one hand, and the domestic price pass-through from fuel prices to wages (indirect effect) on the other. The impulse-response function of the vector autoregressive (VAR) technique was used to examine the direct effects of fuel prices on the inflation rate and wages, as well as the indirect effect from fuel prices to wages in South Africa. Quarterly data from 2001Q1 to 2018Q2 were utilised. The empirical results showed that inflation and wages responded positively to the shock in the fuel price. The results indicated that although wages increased because of a shock in the inflation rate—depicting the pass-through from the fuel price—the direct impact of the fuel price on wages is more significant. This study departs from the usual studies that examine the exchange rate pass-through of international oil prices to the inflation rate. The conclusion drawn was that the pass-through effect was not as strong and did not contribute as much as the direct effect of the fuel price. Therefore, the indirect shock of fuel prices, via inflation, does not so much cause an increase in wages as its direct shock. The study concludes with policy recommendations.

Keywords: fuel prices; impulse-response function; inflation rate; South Africa; wages; vector autoregressive function (VAR)
Introduction

South Africa, like other developing economies, has not been spared the effects of international oil price shocks. There have been considerable debate and studies on the exchange rate pass-through of the oil price in South Africa, whereby the exchange rate responds to oil price shocks and the inflation rate, in turn, responds to changes in the exchange rate. However, the magnitude of the effect of the oil price shock depends on the degree of dependence of the importing country (Nkomo 2006). Crude oil is one of the largest commodity imports by South Africa, making up approximately 11% of its total imports (South African Market Insights 2018b). This shows that South Africa is highly dependent on crude oil, having imported over eight billion kilograms of crude oil, to the value of over R54 billion, in the first five months of 2018 alone (South African Market Insights 2018b).

However, one of the main impacts of oil price changes is on the price level inflation rate. Achieving a low inflation rate, which is a measure of price stability, is important as a main policy objective targeted by monetary policy makers. Changes in oil prices also influence the economic performance via wages. This is because rising oil prices, which lead to increased fuel prices, affect everyone, since the prices of consumer goods and services purchased move in the same direction as the fuel prices. The increased fuel prices, therefore, reflect the underlying increases in production costs (StatsSA 2018). This shows that the increase in fuel prices does not only affect the inflation rate, but further filters through to real wages. It is, therefore, important to examine the extent to which the shock in fuel prices affects the economy via the inflation rate and the wage rate. There is little doubt that increases in fuel prices have a significant impact on the South African economic outlook.

Persistent oil price shocks could result in severe macroeconomic implications, causing challenges in policy making. Since the 1970s, oil price shocks have been blamed for higher inflation and changes in monetary policy, among others (Kilian 2014). While fuel prices are an essential part of price levels, wage bargaining depicts a general pay increase in the price or price expectations. Thus, the aim of the paper is to investigate the direct effect of fuel prices on the inflation rate and wages in South Africa, on the one hand, and the domestic price pass-through from fuel prices to wages (indirect effect) on the other. The purpose of the present study is to contribute to the ongoing research about the effect of the fuel price on the economy.

The objectives of the study are: 1) to examine the responsiveness of the inflation rate and wages to an unexpected fuel price shock; and 2) to study the effect of an unexpected inflation rate shift on domestic wages, using the impulse-response function of the vector autoregressive (VAR) technique. The study first provides the stylised facts on South Africa’s oil sub-sector, followed by a review of the literature. The methodology is presented and analyses of the results are provided. The article concludes with a summary and recommendations.
Stylised Facts on South Africa’s Oil Sub-sector

South Africa imports the majority of its crude oil from three countries, namely Saudi Arabia, Nigeria and Angola, making up about 84% of its total crude oil imports, where different countries charge different amounts. Figure 1 shows the 10 largest countries from where South Africa imported crude oil, from January 2010 to June 2017.

Figure 1: Contributors to South Africa’s crude oil imports

Source: South African Market Insights (2018b)

The average increase in the first five months of 2018 compared to the first five months of 2017 was 11.3%. The fuel (pump) prices, which are calculated on the basis of the exchange rate, world crude oil prices, costs of shipping, taxes and fuel levies, have been increasing over the years. In South Africa, fuel levies are in the form of two main indirect taxes, namely the general fuel levy and the road accident fund (RAF) levy. These two levies constitute about 35% of the petrol price per litre. The petrol price is reviewed and adjusted every month based on the rand/US$ exchange rate and international petrol prices. The study, therefore, looks at the following: \( \text{fuel price} = \text{oil price} + \text{the four components below} \).

According to the Automobile Association (2018), the fuel price comprises the following four main components: the general fuel levy, which is the tax charged on every litre of petrol sold; the money collected from the RAF levy is a portion of every litre of petrol sold, and is used to compensate the victims of road accidents. The other two components are the basic fuel price (BFP), which comprises shipping costs, insurance, storage and the costs to harbour facilities when off-loading petroleum products into storage.
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(wharfage), and lastly other costs which include transport costs, customs and excise duties, retail margins paid to fuel station owners and secondary storage costs. The price of a litre of petrol can either increase or decrease, depending on whether the price of any of these elements increases or decreases. According to South African Market Insights (2018a), 50% of the petrol price comprises taxes, levies, transport costs and the profit margins of wholesalers and retailers.

Therefore, as international oil prices increase, the resultant rippling effect on the South African economy is vast, given the additional components responsible for the increase in the domestic price. The increased general fuel and RAF levies in April 2019 and additional tax (carbon tax) on the fuel price could see further increases in inflation and thus wage rates. Given that international oil prices are driven by demand and supply, while fuel prices are calculated on the basis of the exchange rate, among other factors, this study therefore departs from the usual studies that examine the exchange rate pass-through of international oil prices to the inflation rate.

Figure 2 shows the trend in the international oil price and the fuel (pump) price in South Africa. It is clear that they both have the same movement, in the sense that as international oil prices increase (or decrease), so do the fuel prices. Although this shows that either of the two could be used, in this study the fuel price was utilised as it has more bearing on consumers’ behaviour and reactions. For instance, consumers react more to the announcement of the increase in fuel (pump) price, compared to when there is an increase in international oil prices.

**Figure 2: International oil and fuel price trend (2001–2018)**

Graph compiled by authors  
**Source:** SARB Database (2018) and Department of Energy Database (2018)

Although the fuel price in South Africa is linked to the international oil price, while the fuel levies, taxes, and so forth do not change significantly in the short run, they do
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constitute a major portion of the total fuel price. Hence the domestic fuel price is a better reflection of the impact on domestic wages and inflation rates than the international oil price.

Literature Review

Theoretical underpinnings of changes in the oil price rely on both its direct effects—made up of the aggregate demand shock and the aggregate supply shock—and the indirect effects (Asghar and Naveed 2015). While the demand shock is the reduction in the purchasing power of domestic households, the supply shock is the increase in the production cost of domestic output. On the other hand, an oil price shock has an impact on wages, with the possibility of demand for higher wages by unions in order to offset the increases in prices resulting from the oil price shocks (Kilian 2014). Many theoretical models generally predicted that oil-price shocks cause wages and prices to increase, while real output declines (Bruno 1982; Burbidge and Harrison 1984).

A number of empirical studies have been conducted on the oil price pass-through to inflation in both developed and developing countries to support the theoretical findings. However, no study has observed the fuel price pass-through to real wages in particular. Chen (2009), for instance, investigated the oil price pass-through to the inflation of 19 industrialised countries, using a state-space approach to estimate time-varying oil price pass-through coefficients. The study found that there was a declining pass-through for almost all of the countries included in the study. According to Chen (2009), the variables that explain the decline in oil pass-through are the appreciation of the domestic currency, a more active monetary policy in response to inflation, and a higher degree of trade openness.

Likewise, Álvarez et al. (2011) assessed the impact of oil price changes on consumer price inflation within the Spanish and Euro area, using the dynamic stochastic general equilibrium (DSGE) model. The study found that the inflationary effect of oil price changes in both economies was limited, even though crude oil price fluctuations were a major driver of inflation variability. Using the dynamic error correction model (ECM), Abounoori, Nazarian, and Amiri (2014) investigated the nature and causes of the oil price pass-through to inflation in the short and long run in Iran, using monthly data from 2003M3 to 2013M3. The study found that the oil price pass-through to inflation in both the short and long run was positive.

Furthermore, Sek, Teoa, and Wong (2015) studied the effects of oil price changes in high and low oil dependency country groups, using data from 1980 to 2010. The pass-through equation was modelled in an autoregressive distributed lag (ARDL) format and estimated using the pooled mean group method. The results showed a direct effect of the oil price change on domestic inflation in the low oil-dependency group, and an indirect effect in the high oil-dependency countries through changes in the exporters’ production cost. They further stated that the higher production cost would then pass through to domestic price levels and indirectly increase domestic inflation.
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Long and Liang (2018) utilised both the autoregressive distributed lag (ARDL) and nonlinear and asymmetric autoregressive distributed lag (NARDL) models to investigate the pass-through effects of the crude oil price on China’s producer prices index (PPI) and consumer prices index (CPI). Their results indicated that the impact of global oil price fluctuations on PPI and CPI were asymmetrical in the long run. The study found that the long-term impacts of the rise in global oil prices on PPI and CPI were greater than the global oil price decline on PPI and CPI.

Živkov, Đurašković, and Manić (2019) investigated the effects of oil price changes on consumer price inflation in 11 Central and Eastern European countries, using the wavelet-based Markov switching approach in order to distinguish between the effects at different time horizons. They found that the transmission of oil price changes to inflation was relatively low in these countries because an increase in the oil price of 100% was followed by a rise in inflation of one to six percentage points. Živkov et al. (2019) also found that when oil shocks were transmitted towards inflation, the exchange rate was not a significant factor, except when high depreciation occurred. The countries that were found to have experienced the highest and most consistent pass-through in the study were Slovakia and Bulgaria, and this could be due to the fact that these countries have some of the highest oil import/GDP ratios.

However, Peneva and Rudd (2017) assessed how the pass-through of labour costs to price inflation has evolved over time, using a time-varying parameter/stochastic volatility VAR framework. Using USA data, they found little evidence that independent movements in labour costs have had a material effect on price inflation in recent years. By identifying the inflation regimes in Canada after 1961, Faroque and Minor (2009) found that both in and out-of-sample evidence suggested that wage growth exerts an influence on inflation only during a high-inflation regime, but inflation exerts a more systematic and quantitatively stronger influence on wage growth regardless of the prevailing inflation regime. These results did not support the view of cost-push inflation.

Conflitti and Luciani (2017) estimated the oil price pass-through to consumer prices of the USA and the Euro area by first estimating a dynamic factor model on a panel of disaggregated price indicators, and then using VAR techniques to estimate the pass-through. The results showed that the oil price passes through core inflation only via its effect on the whole economy, and although the pass-through is estimated to be small, it is statistically different from zero and is also long lasting.

Over the years, there have been numerous attempts to investigate the impact of oil price shocks on a number of macroeconomic variables in South Africa. Some of the studies include the work of Kohler (2006), Nkomo (2006), Hollander, Gupta, and Wohar (2018) and Tshepo (2015). For instance, Nkomo (2006) examined crude oil price movements and their impact on the South African economy. The study found that the country has been shielded from much of the negative impacts of crude oil price increases because of the strong exchange rate between the rand and US dollar, but is still highly dependent.
on and vulnerable to external sources of oil supply and to increases in international oil prices.

Kin and Courage (2014) investigated the impact of oil prices on the nominal exchange rate in South Africa, using the generalised autoregressive conditional heteroscedasticity (GARCH) technique and data from 1994 and 2012. The results indicated that oil prices had a significant impact on nominal exchange rates and an increase in oil prices led to a depreciation of the rand exchange rate.

Dadam and Viegi (2015) analysed the effect of the labour market on the conduct of monetary policy in South Africa by observing the effect of wage inflation on inflation expectations. The study used different econometric techniques such as ordinary least squares (OLS), the dynamic stochastic general equilibrium model, and line graphs. Many specifications of the new Keynesian wage Phillips curve and the reduced form traditional Phillips curve were estimated. Their results showed a strong correlation between wage inflation and the inflation expectations of the trade unions. The study thus showed that regulating inflation expectations can control changes in wages.

By contrast, Sibanda, Hove, and Murwirapachena (2015) examined the impact of the oil price and exchange rate on inflation expectations, rather than on the inflation rate. Their study also did not consider the effect of the oil price change on wage rates. They used monthly time-series data spanning 2002M7 to 2013M3, and the vector autoregressive model. Their findings indicated that both oil prices and exchange rates had strong and significantly positive impacts on inflation expectations. Their study further examined the effect of inflation expectations on inflation rates and found a significant one.

More recently, using the vector autoregression and the vector error correction mechanism methods, Sedick (2016) found that the movement in Brent oil prices had a relatively insignificant impact on the movement of the South African rand on a monthly basis. Nzimande and Msomi (2016) examined the relationship between oil prices and economic activity in South Africa, using the asymmetric cointegration approach and data from 1966 to 2014. The results of the study supported the findings of other studies, which reported that economic activity responds asymmetrically to oil price shocks, such that an increase in energy prices does not affect economic activity, whereas a decrease in oil prices stimulates economic activity.

Furthermore, Chisadza et al. (2013) investigated the impact of the oil supply and demand shocks on the South African economy, using a sign restriction-based structural vector autoregressive (VAR) model and quarterly data for the period 1975Q1 to 2011Q2. The results of the study indicated that oil supply shock had a short-lived significant impact only on the inflation rate, while the impact on the other macroeconomic variables was statistically insignificant. Chisadza et al. (2013) further stated that the inflation rate and the real exchange rate reacted negatively to an oil-
specific demand shock, while output was positively related to unanticipated changes in oil price due to speculations.

Lastly, using the auto regressive distributed lag (ARDL) method of co-integration, Rangasamy (2017) analysed the impact of petrol price movements on inflation outcomes in South Africa after the mid-1970s. The results indicated that petrol prices had a direct impact on the prices of other (non-petrol) commodities in the economy and that increases in the petrol price had a significant bearing on inflation outcomes in South Africa.

The studies reviewed show the impact of an oil price shock on inflation via the exchange rate in different countries, using various methodologies. Most of the studies centred around the oil price pass-through to inflation and did not observe the oil price nor fuel price pass-through to wages. Given the above empirical reviews, this study aims to add to the existing body of knowledge by examining the effects of the fuel price shock directly on the inflation rate and wage inflation in South Africa, using the impulse-response function.

Methodology

Data Source and Model Specification

The domestic fuel price (FUEL) is the price per litre at the pumps. This was obtained from the Department of Energy database. The input cost variable, labour cost (WAGES) was measured as the remuneration per worker in the non-agricultural sector, at constant prices (2010 = 100). This was obtained from the South African Reserve Bank (SARB) quarterly bulletin via the Econostat database. The consumer price index, measure of inflation rate (INFL) variable was obtained from the International Monetary Fund (IMF/IFS) database. Quarterly data, spanning over the period 2001Q1 to 2018Q2, were adopted. These periods covered the post-1970s oil crisis, the 1990s and mainly the 2000s oil price shock. The oil price shock of 1990, referred to as “mini oil-shock,” which was the result of increased oil prices and lasted for nine months, was excluded from the data. However, the oil crisis in the 2000s witnessed different surges of price shock. The choice of variables was guided by the objectives of the study and their effects on the South African economy.

The model to be estimated is as follows:

\[
\bar{W}_t = \alpha_0 + a(L)FUEL_t + b(L)WAGES_t + c(L)INFL_t + \varepsilon_t \quad \ldots \{1\}
\]

where \(\bar{W}_t\) is the vector of dependent variables, L is the lag operator and \(\varepsilon_t\) is the vector of error terms. All the variables are as defined above.

Given South Africa’s dependence on imported crude oil, it is thus exposed to increases in imported inflation as well as possible increases in input cost (real wages in this case).
The recent increases in the international oil prices tend to affect domestic prices as a result of an increase in the cost of production and transport costs of the importing country. While fuel price (FUEL) has a direct impact on inflation dynamics and real wages, on the one hand, it has an indirect effect on real wages via inflation rate, on the other. It is, therefore, necessary to investigate the effects of the direct fuel price shock on real wages and inflation dynamics in South Africa. In addition, the indirect effect of the fuel price shock is equally important by way of a possible price pass-through to wages. This shows how labour reacts to inflation dynamics resulting from the fuel price shock.

Estimation Techniques

Economic time series are generally non-stationary (Johansen and Juselius 1990; Hjalmarsson and Osterholm 2007). Hence, before variables can be subjected to cointegration tests, the order of integration must be established. For this purpose, the study utilises the augmented Dickey-Fuller (ADF), and the Phillips-Perron (PP) tests. If the variables are integrated of order zero, I(0), then a simple vector autoregressive (VAR) technique is used. However, if they are integrated of different orders, their long-run relationship is observed using the appropriate technique.

The estimation procedure is as follows: following the tests of stationarity, depending on the order of integration, the test for cointegration will be applied. The VAR model will be estimated in its reduced form, followed by the diagnostic test of stability. Finally, the impulse-response function, which is the crux of the study, will thus be carried out. Although the ultimate goal of this study is to observe the effect of the shock in domestic fuel price on wages via inflation rate variables, the analysis was started in a VAR system. VAR is a more appropriate technique because it captures the dynamic nature of multiple time-series and provides reliable policy analysis. By applying a VAR technique, which assumes that all the variables are endogenous, the system of equations is such that each variable is determined by other variables (Sims 1980) (equation 2).

Hence the multivariate linear simultaneous VAR system for domestic price pass-through can be written as follows:

\[ FUEL_i = \alpha_0 + \sum_{i=1}^{n_1} \alpha_i FUEL_{t-i} + \sum_{i=1}^{n_2} \alpha_{2i} WAGES_{t-i} + \sum_{i=1}^{n_3} \alpha_{3i} INFL_{t-i} + \mu_i \]

\[ INFL_i = \alpha_0 + \sum_{i=1}^{n_1} \alpha_i FUEL_{t-i} + \sum_{i=1}^{n_2} \alpha_{2i} WAGES_{t-i} + \sum_{i=1}^{n_3} \alpha_{3i} INFL_{t-i} + \mu_i \]

\[ WAGES_i = \alpha_0 + \sum_{i=1}^{n_1} \alpha_i FUEL_{t-i} + \sum_{i=1}^{n_2} \alpha_{2i} WAGES_{t-i} + \sum_{i=1}^{n_3} \alpha_{3i} INFL_{t-i} + \mu_i \cdots \{2\} \]

Each of the variables in the VAR system is endogenous, except for the constant term.
The lag length at which the VAR will be estimated was first determined, using the Akaike information criterion (AIC), and the stability test was performed on the estimated VAR. The stability of the VAR roots is the main diagnostic test that has to be passed before continuing the estimation in order to produce reliable, stable and valid results. If the VAR is not stable, results such as the impulse-response and the variance decomposition will be invalid.

The impulse-response function (IRF) and the variance decomposition (VD) within the VAR system use the innovation accounting technique to analyse the dynamic relationship among variables. The IRF gives the interaction and the response between variables, while the VD provides the contribution of each variable to the shock in other variables (Morales 2003). Thus, this study observes the innovation in the selected macroeconomic variables, given a negative shock in the domestic fuel price. The IRF traces the effect of one standard deviation negative shock in the fuel price on both the current and future values of the inflation rate and via the inflation rate to wages (equation 2) in the system. The IRF observes how a given variable responds over a specified period of time (the time path), resulting from a one-unit negative shock in the fuel price variable.

The importance of an IRF technique is its applicability to policy analysis because it assists policy makers to observe how macroeconomic variables respond to shocks in a given variable, as well as the duration of the effects of the shocks before they die out (if at all). Consequently, the results obtained from these methods should assist policy makers in observing the variables that are mostly affected as a result of a shock in the domestic fuel price in South Africa (IRF) and the variables that contribute more to the decomposition of the fuel price (VD).

This study used the generalised impulses decomposition method and Monte Carlo response standard errors of the IRF. Monte Carlo response standard errors decrease the variability in the mean estimate. The generalised impulses decomposition method was chosen because, unlike the Cholesky decomposition method of IRF—whose responses change drastically as the ordering of the variables changes—the generalised impulses decomposition method constructs an orthogonal set of innovations that does not depend on the ordering of the VAR.

Results

The simple correlation matrix shows the relationship between each of the variables, fuel price, inflation rate and labour cost. There is a positive but statistically insignificant correlation between the fuel price and inflation rate. The low positive correlation of about 13% indicates that the fuel price and inflation rate are positively related, as expected. However, labour cost shows a positive and highly statistically significant correlation with the fuel price. The high positive correlation shows that fuel price and labour cost have a correlation of 95%, which supports the a priori expectation of a
positive relationship. Furthermore, the inflation rate and labour cost have a negative correlation of 1%. Table 1 presents the correlation results.

**Table 1: Correlation matrix: South Africa, 1990–2018**

|       | FUEL | INFL  | WAGES |
|-------|------|-------|-------|
| FUEL  | 1    |       |       |
| INFL  | 0.126| 1     |       |
| WAGES | 0.952***| -0.010| 1     |

Author’s analysis
***1%, **5%, *10%

Since all the variables are stationary after the first difference, the study proceeded with the cointegration test. The stationarity test results are presented in table 2.

**Table 2: Stationarity test results**

| Variables | LEVELS | FIRST DIFFERENCE | Conclusion |
|-----------|--------|------------------|------------|
|           | ADF    | PP               | ADF        | PP         |            |
| FUEL      | -0.486 | 0.030            | -7.865***  | 9.695***   | I(1)       |
| INFL      | -2.948**| -2.873*         | -7.451***  | -4.366***  | I(1)       |
| WAGES     | -1.138 | -1.138           | -10.236*** | -10.625*** | I(1)       |

Critical values: [1% -3.537; 5% -2.908; 10% -2.591]; ***1%, **5%, *10%

Since all the variables are I(1), the study went further to test for cointegration among the variables, using the Johansen cointegration technique. However, before testing for cointegration, the study estimated the VAR at the lag length of 2, chosen by Akaike information criterion. The VAR stability test was performed and the VAR roots were found to lie inside the unit circle, which confirms that that VAR is stable (table A1, appendix A). The Johansen test for cointegration found that there is no cointegration among the variables at the chosen lag of 2 (table A2, appendix A). This is not surprising, given the weight of fuel price in the consumer price inflation of about 25.3% (StatsSA 2018).

The study then observed the effect of the one standard deviation shock in the fuel price on the selected macroeconomic variables, and in order to observe the price pass-through, the one standard deviation shock in inflation rate on the wage rate was examined, using the impulse-response function method. Figure 3 presents the impulse-response results.
Figure 3: Price pass-through of fuel prices over a two-year period

The results show how each of the variables responded to the shock in another variable over a period of eight quarters (2 years). The Monte Carlo response standard errors show the plus/minus two standard error bands around the impulse response. Although the variables responded as expected, the importance of the impulse-response function goes beyond variables merely meeting expectations; it is also to trace out the duration of a one standard deviation shock in the fuel prices on inflation rate and its pass-through to wages. Given a standard deviation shock in the fuel price, the results show that the inflation rate initially increased after the second quarter up to the fifth quarter before it started to decline.

Generally, in the first year of the shock in the fuel price, that is, in the first four quarters, the inflation rate responded by increasing gradually, then started to decline steadily in the second year (from the sixth to the eighth quarter), and settled at a point lower than it was at the onset of the shock. However, there is generally a positive relationship between the inflation rate and fuel price during the two-year period under investigation. Meanwhile, the wages increased steadily in response to a shock in fuel price, from the first to the eighth quarter. Although the relationship between wages and fuel price started out as negative, it largely depicted a positive relationship during the entire two years.
The price pass-through of the fuel price to wages shows a somewhat interesting picture. While wages responded immediately to a direct shock in fuel prices, they responded more quickly to the shock via inflation rate, but the response was slow. Given a shock in inflation rate, wages increased more steadily, although they initially declined from the first to the third quarter; then they picked up and continued to increase until the eighth quarter. There was a negative relationship between wages and inflation rate in the first six quarters as a result of fuel price shock. This changed to a positive relationship after the sixth quarter and continued to remain positive.

While the direct response of wages to the shock in fuel price increased over the entire two-year period, the pass-through effect showed an initial decline and only started increasing after the third quarter. Also, wages responded positively to the shock in the fuel price immediately, and continued to be positively related throughout the eight quarters. Meanwhile, wages initially responded negatively to the shock via the inflation rate before indicating a positive relationship after a year. This confirms the earlier statement that the fuel price has a greater bearing on the behaviour and reaction of consumers, because they react more to the announcement of the increase in fuel (pump) price. As soon as an increase in the fuel price is announced, consumers expect inflation to increase and they demand higher wages.

The variance decomposition of wages shows the variables that explain their decomposition. The variance decomposition results are presented in table 3.
Table 3: Two-year (8 quarters) contributions of fuel price and inflation rate to the innovations in wages

| Period | S.E.     | FUEL | INFL  | WAGES  |
|--------|----------|------|-------|--------|
| 1      | 0.965488 | 0.000000 | 0.819129 | 99.18087 |
|        | (0.00000) | (3.12355) | (3.12355) |        |
| 2      | 1.721546 | 0.678757 | 1.713449 | 97.60779 |
|        | (1.76671) | (4.03802) | (4.59568) |        |
| 3      | 2.277605 | 1.401763 | 2.095174 | 96.50306 |
|        | (3.21015) | (5.01188) | (6.12726) |        |
| 4      | 2.610894 | 1.935739 | 2.082168 | 95.98209 |
|        | (4.78380) | (5.72334) | (7.54808) |        |
| 5      | 2.769587 | 2.257680 | 1.847413 | 95.89491 |
|        | (6.03251) | (6.21367) | (8.58529) |        |
| 6      | 2.829954 | 2.456634 | 1.600119 | 95.94325 |
|        | (7.00286) | (6.69166) | (9.41637) |        |
| 7      | 2.859396 | 2.605421 | 1.453419 | 95.94116 |
|        | (7.77159) | (7.28125) | (10.1874) |        |
| 8      | 2.893688 | 2.752145 | 1.418425 | 95.82943 |
|        | (8.41673) | (7.95971) | (10.9458) |        |

Cholesky ordering: FUEL INFL WAGES

Standard errors: Monte Carlo (100 repetitions)

Author’s analysis

While the contribution from the inflation rate showed a clear contribution to the innovations in wages in the first year of the period under consideration, the fuel price made a stronger, significant and considerable contribution during the second year. This shows that although wages increased as a result of a shock in the inflation rate, depicting the pass-through from the fuel price (this was the case in the first four quarters), the fuel price had a more significant and direct effect on wages, starting from the second year. The decline in inflation contribution to the innovation in wages thereby showing no strong indirect effect of fuel price to wage, could be due to the good management of inflation targeting framework in the country. The longer period of the contributions of these variables is shown in the appendix (table A3, appendix A).

Therefore, the demand for increased wages is a direct result of increased fuel prices. Consumers respond more quickly to the increased fuel price than they react to the rise in the inflation rate, to demand higher wages. This also confirms the correlation matrix (table 1), where there was more than a 90% correlation between the fuel price and wages. The study passed the diagnostic tests.

Summary and Conclusion

The aim of this study was twofold, namely to: 1) observe the direct responsiveness of the inflation rate and wages to the domestic fuel price shock; and 2) investigate the pass-
through effect of the fuel price via the inflation rate to domestic wages. This analysis was conducted using quarterly data over the period 2001Q1 to 2018Q2 and adopting the impulse-response function and variance decomposition techniques. The study found that both the inflation rate and wages responded to the shock in the fuel price as anticipated. The study further showed that the higher percentage variance in wages was the result of the direct change in the fuel price, rather than via the inflation rate.

There was a gradual increase in the inflation rate as a result of an unexpected shock in the domestic fuel price, but this increase was short lived as the inflation rate gradually declined after about the fifth quarter. This shows that at the onset of the oil price shock, the inflation rate increased, but it declined after about a year and a half. Thus, monetary policy that shows the time lag of about 24 months is applicable to South Africa, and the monetary authorities should therefore not raise interest rates every two months, as this will amplify the economic contraction. On the other hand, given an unexpected shock in the fuel price, wages responded directly by increasing steadily for more than a year, while their response to the fuel price shock via the inflation rate was an ephemeral increase, starting with a negative relationship up to the sixth quarter before changing to a positive one, while still increasing.

The results also showed that while wages reacted to the increase in the inflation rate, they responded more directly to the changes in fuel prices. This indicates that although wages increased as a result of a shock in the inflation rate, depicting the pass-through from the fuel price, the fuel price had a more significant and direct effect on wages. While this study does not conclude that there was no price pass-through of the fuel price to wages, it emphasises that the pass-through effect was not as strong and did not contribute as much as the direct effect from the fuel price. The reason could be the proper management of inflation, and the inflation targeting framework of the monetary policy committee of the South African Reserve Bank. The study, therefore, suggests that seeing the huge and stronger effect of fuel prices directly on wages, the government could consider lowering the fuel levies, tax, and so forth.

While many studies have considered the exchange rate pass-through or the effects of the oil price on the inflation rate, it is imperative to understand the effects of the oil price and fuel prices on the macro economy, especially in a country like South Africa, which faces continuous increases in fuel prices. Further research includes considering the use of different techniques, such as inflation expectations instead of inflation rate, or examining the effects of the fuel levies as a component of fuel price on household consumption expenditure and other macroeconomic variables.

References
Abounoori, A. A., R. Nazarian, and A. Amiri, A. 2014. “Oil Price Pass-through into Domestic Inflation: The Case of Iran.” International Journal of Energy Economics and Policy 4 (4): 662–669.
Álvarez, L. J., S. Hurtado, I. Sánchez, and C. Thomas. 2011. “The Impact of Oil Price Changes on Spanish and Euro Area Consumer Price Inflation.” *Economic Modelling* 28 (1–2): 422–431. https://doi.org/10.1016/j.econmod.2010.08.006.

Asghar, N., and T. A. Naveed. 2015. “Pass-through of World Oil Prices to Inflation: A Time Series Analysis of Pakistan.” *Pakistan Economic and Social Review* 53 (2): 269–284.

Automobile Association. 2018. “Petrol Price Breakdown.” Accessed April 8, 2019. https://www.aa.co.za/insights/petrol-price-breakdown.

Bruno, M. 1982. “Adjustment and Structural Change under Supply Shocks.” *Scandinavian Journal of Economics* 84: 199–221. https://doi.org/10.2307/3439635.

Burbidge, J., and A. Harrison. 1984. “Testing for the Effects of Oil-price Rises Using Vector Autoregression.” *International Economic Review* 25 (2): 459–484. https://doi.org/10.2307/2526209.

Chen, S. 2009. “Oil Price Pass-through into Inflation.” *Energy Economics* 31 (1): 126–133. https://doi.org/10.1016/j.eneco.2008.08.006.

Chisadza, C., J. Dlamini, R. Gupta, and M. P. Modise. 2013. “The Impact of Oil Shocks on the South African Economy.” Department of Economics, University of Pretoria, Working Paper Series No. 201311.

Conflitti, C., and M. Luciani. 2017. “Oil Price Pass-through into Core Inflation.” Board of Governors of the Federal Reserve System, Finance and Economics Discussion Series 2017–085. https://doi.org/10.17016/FEDS.2017.085.

Dadam, V., and N. Viegi. 2015. “Labour Market and Monetary Policy in South Africa.” *ERSA Working Paper*, No. 551, September.

Department of Energy Database. 2018.

Faroque, A., and R. Minor. 2009. “Inflation Regimes and the Stability of the Pass-through of Wages to Consumer Prices in Canada.” *Applied Economics* 41 (8): 1003–1017. https://doi.org/10.1080/00036840601019133.

Hjalmarmsson, E., and P. Osterholm. 2007. “Testing for Cointegration using the Johansen Methodology when Variables are Near-integrated.” International Monetary Fund Working Paper, WP/07/141. https://doi.org/10.5089/9781451867053.001.

Hollander, H., R. Gupta, and M. E. Wohar. 2018. “The Impact of Oil Shocks in a Small Open Economy: New-Keynesian Dynamic Stochastic General Equilibrium Model for an Oil-Importing Country: The Case of South Africa.” *Emerging Markets Finance and Trade*, 2–26. https://doi.org/10.1080/1540496X.2018.1474346.
Leshoro, Maluleke

Johansen, S., and K. Juselius. 1990. “Maximum Likelihood Estimation and Inference on Cointegration: With Applications to the Demand for Money.” *Oxford Bulletin of Economics and Statistics* 52 (2): 169–210. https://doi.org/10.1111/j.1468-0084.1990.mp52002003.x.

Kilian, L. 2014. “Oil Price Shocks: Causes and Consequences.” *Annual Review of Resources Economics* 6: 133–154. doi: 10.1146/annurev-resource-083013-114701. https://doi.org/10.1146/annurev-resource-083013-114701.

Kin, S., and M. Courage. 2014. “The Impact of Oil Prices on the Exchange Rate in South Africa.” *Journal of Economics* 5 (2): 193–199. https://doi.org/10.1080/09765239.2014.11884996.

Kohler, M. 2006. “The Economic Impact of Rising Energy Prices: A Constraint on South Africa’s Growth and Poverty Reduction Opportunities.” Paper prepared for the TIPS/DPRU Forum, Johannesburg, 18 to 20 October.

Long, S., and J. Liang. 2018. “Asymmetric and Nonlinear Pass-through of Global Crude Oil Price to China’s PPI and CPI Inflation.” *Economic Research-Ekonomiska Istraživanja* 31 (1): 240–251. https://doi.org/10.1080/1331677X.2018.1429292.

Morales, M. 2003. “Dynamic Interaction for Andean Countries: Evidence from VAR Approach.” *Economic Commission for Latin America and the Caribbean, ECLAC*. REDIMA Work-group Meeting on Modelling Macroeconomic Coordination in the Andean Community, Santiago, Chile, October, 22–20 October.

Nkomo, J. C. 2006. “The Impact of Higher Oil Prices on Southern African Countries.” *Journal of Energy Research in Southern Africa* 17 (1): 10–17. https://doi.org/10.17159/2413-3051/2006/v17i1a3373.

Nzimande, N., and S. Msomi. 2016. “Oil Price Shocks and Economic Activity: The Asymmetric Cointegration Approach in South Africa.” SAEF Working Paper 01/05. MWG01, 1–15.

Peneva, E. V., and J. B. Rudd. 2017. “The Pass-through of Labor Costs to Price Inflation.” *Journal of Money, Credit and Banking* 49 (8): 1777–1802. https://doi.org/10.1111/jmcb.12449.

Rangasamy, L. 2017. “The Impact of Petrol Price Movements on South African Inflation.” *Journal of Energy in Southern Africa* 28 (1): 120–132. https://doi.org/10.17159/2413-3051/2017/v28i1a1597.

Sedick, A. 2016. “An Analysis of the Impact of Crude Oil Price Shocks on the Exchange Rate in South Africa.” Unpublished master’s dissertation, University of the Western Cape, Cape Town.
Leshoro, Maluleke

Sek, S. K., X. O. Teoa, and Y. N. Wong. 2015. “A Comparative Study on the Effects of Oil Price Changes on Inflation.” Procedia Economics and Finance 26: 630–636. https://doi.org/10.1016/S2212-5671(15)00800-X.

Sibanda, K., P. Hove, and G. Murwirapachena. 2015. “Oil Prices, Exchange Rates and Inflation Expectations in South Africa.” International Business and Economics Research Journal 14 (4): 587–602. https://doi.org/10.19030/iber.v14i4.9351.

Sims, C. A. 1980. “Macroeconomics and Reality.” Econometrica 48: 1–49. https://doi.org/10.2307/1912017.

South African Market Insights. 2018a. “Latest Fuel Prices in South Africa.” Accessed April 8, 2019. https://www.southafricanmi.com/latest-fuel-prices-in-sa-5jul2018.html.

South African Market Insights 2018b. “South Africa’s Crude Oil Imports from January 2018 to May 2018.” Accessed April 8, 2019. https://www.southafricanmi.com/south-africas-crude-oil-imports-6jul2018.html.

South African Reserve Bank (SARB) Database. 2018.

Statistics South Africa (StatsSA). 2018. “Fuel Inflation Remains in Double-digit Territory as Prices Rise.” Accessed April 8, 2020. http://www.statssa.gov.za/?p=11558.

Tshepo, M. 2015. “Analysing the Pass-through Effects of Oil Prices on Inflation in South Africa: Granger-Causality Approach.” Biennial Economic Society of South Africa Conference, Cape Town, South Africa.

Živkov, D., J. Đurašković, and S. Manić. 2019. “How Do Oil Price Changes Affect Inflation in Central and Eastern European Countries? A Wavelet-based Markov Switching Approach.” Baltic Journal of Economics 19 (1): 84–104. https://doi.org/10.1080/1406099X.2018.1562011.
Appendix A: Tables

Table A1: VAR roots

| Root          | Modulus  |
|---------------|----------|
| 0.990086      | 0.990086 |
| 0.767547 - 0.367551i | 0.851013 |
| 0.767547 + 0.367551i | 0.851013 |
| 0.538708      | 0.538708 |
| -0.219270     | 0.219270 |
| 0.159395      | 0.159395 |

No root lies outside the unit circle.
VAR satisfies the stability condition.

Table A2: Johansen cointegration

Sample (adjusted): 2001Q4 2018Q2
Included observations: 67 after adjustments
Trend assumption: Linear deterministic trend
Series: FUEL INFL WAGES
Lags interval (in first differences): 1 to 2

Unrestricted Cointegration Rank Test (Trace)

| Hypothesised                   | Trace       | Critical Value | Prob.** |
|--------------------------------|-------------|----------------|---------|
| None                           | 0.193419    | 29.79707       | 0.2758  |
| At most 1                      | 0.095706    | 15.49471       | 0.4635  |
| At most 2                      | 0.018890    | 3.841466       | 0.2583  |

Trace test indicates no cointegration at the 0.05 level
* denotes rejection of the hypothesis at the 0.05 level
**MacKinnon-Haug-Michelis (1999) p-values

Unrestricted Cointegration Rank Test (Maximum Eigenvalue)

| Hypothesised                   | Max-Eigen   | Critical Value | Prob.** |
|--------------------------------|-------------|----------------|---------|
| None                           | 0.193419    | 21.13162       | 0.3330  |
| At most 1                      | 0.095706    | 14.26460       | 0.5203  |
| At most 2                      | 0.018890    | 3.841466       | 0.2583  |

Max-eigenvalue test indicates no cointegration at the 0.05 level
* denotes rejection of the hypothesis at the 0.05 level
**MacKinnon-Haug-Michelis (1999) p-values
Table A3: Eighteen-quarter contributions of the fuel price and inflation rate to the innovations in wages

| Period | S.E.     | FUEL  | INFL  | WAGES   |
|--------|----------|-------|-------|---------|
| 1      | 0.965488 | 0.000000 | 0.819129 | 99.18087 |
|        | (0.00000) | (3.12355) | (3.12355) |         |
| 2      | 1.721546 | 0.678757 | 1.713449 | 97.60779 |
|        | (1.76671) | (4.03802) | (4.59568) |         |
| 3      | 2.277605 | 1.401763 | 2.095174 | 96.50306 |
|        | (3.21015) | (5.01188) | (6.12726) |         |
| 4      | 2.610894 | 1.935739 | 2.082168 | 95.98209 |
|        | (4.78380) | (5.72334) | (7.54808) |         |
| 5      | 2.769587 | 2.257680 | 1.847413 | 95.89491 |
|        | (6.03251) | (6.21367) | (8.58529) |         |
| 6      | 2.829954 | 2.456634 | 1.600119 | 95.94325 |
|        | (7.00286) | (6.69166) | (9.41637) |         |
| 7      | 2.859396 | 2.605421 | 1.453419 | 95.94116 |
|        | (7.77159) | (7.28125) | (10.1874) |         |
| 8      | 2.893688 | 2.752145 | 1.418425 | 95.82943 |
|        | (8.41673) | (7.95971) | (10.9458) |         |
| 9      | 2.936648 | 2.922549 | 1.447182 | 95.63027 |
|        | (8.98397) | (8.60183) | (11.6550) |         |
| 10     | 2.976943 | 3.125184 | 1.484033 | 95.39078 |
|        | (9.50103) | (9.10753) | (12.2774) |         |
| 11     | 3.005004 | 3.355280 | 1.495221 | 95.14950 |
|        | (9.98202) | (9.44932) | (12.8050) |         |
| 12     | 3.019614 | 3.599261 | 1.473238 | 94.92750 |
|        | (10.4302) | (9.65142) | (13.2499) |         |
| 13     | 3.025642 | 3.840153 | 1.426746 | 94.73310 |
|        | (10.8406) | (9.75244) | (13.6273) |         |
| 14     | 3.028897 | 4.062711 | 1.368515 | 94.56877 |
|        | (11.2066) | (9.78470) | (13.9475) |         |
| 15     | 3.032655 | 4.256815 | 1.308269 | 94.43492 |
|        | (11.5265) | (9.77242) | (14.2178) |         |
| 16     | 3.037230 | 4.418518 | 1.251170 | 94.33031 |
|        | (11.8038) | (9.73686) | (14.4470) |         |
| 17     | 3.041467 | 4.549148 | 1.199329 | 94.25152 |
|        | (12.0446) | (9.69876) | (14.6471) |         |
| 18     | 3.044373 | 4.653439 | 1.153601 | 94.19296 |
|        | (12.2545) | (9.67631) | (14.8310) |         |

Cholesky ordering: FUEL INFL WAGES

Standard errors: Monte Carlo (100 repetitions)

Analysis by authors