Small-scale galaxy clustering in the EAGLE simulation

M. Celeste Artale,1* Susana E. Pedrosa,1 James W. Trayford,2 Tom Theuns,2 Daniel J. Farrow,3 Peder Norberg,2,4 Idit Zehavi,5 Richard G. Bower2,4 and Matthieu Schaller2

1Instituto de Astronomía y Física del Espacio (IAFE,CONICET-UBA), C.C. 67 Suc. 28, C1428ZAA Ciudad de Buenos Aires, Argentina
2Institute for Computational Cosmology, Department of Physics, Durham University, South Road, Durham DH1 3LE, UK
3Max-Planck-Institut für extraterrestrische Physik, Postfach 1312 Giessenbachstrasse, D-85741 Garching, Germany
4Centre for Extragalactic Astronomy, Department of Physics, Durham University, South Road, Durham DH1 3LE, UK
5Department of Astronomy, Case Western Reserve University, 10900 Euclid Avenue, Cleveland, OH 44106, USA

ABSTRACT

We study present-day galaxy clustering in the EAGLE cosmological hydrodynamical simulation. EAGLE’s galaxy formation parameters were calibrated to reproduce the redshift \( z = 0.1 \) galaxy stellar mass function, and the simulation also reproduces galaxy colours well. The simulation volume is too small to correctly sample large-scale fluctuations and we therefore concentrate on scales smaller than a few mega parsecs. We find very good agreement with observed clustering measurements from the Galaxy And Mass Assembly (GAMA) survey, when galaxies are binned by stellar mass, colour or luminosity. However, low-mass red galaxies are clustered too strongly, which is at least partly due to limited numerical resolution. Apart from this limitation, we conclude that EAGLE galaxies inhabit similar dark matter haloes as observed GAMA galaxies, and that the radial distribution of satellite galaxies, as a function of stellar mass and colour, is similar to that observed as well.

Key words: galaxies: evolution – galaxies: formation – galaxies: haloes – galaxies: statistics – cosmology: theory – large-scale structure of Universe.

1 INTRODUCTION

The spatial distribution of galaxies provides a powerful way to probe both cosmology and galaxy formation. Galaxy clustering measurements on scales where density fluctuations are only mildly non-linear, combined with other cosmological data sets such as cosmic microwave background measurements, put impressively tight constraints on cosmological parameters (e.g. Hinshaw et al. 2013; Planck Collaboration XIII 2016). In addition, the detection of baryon acoustic oscillations in the clustering of galaxies (e.g. Cole et al. 2005; Eisenstein et al. 2005) opened up the way to quantify the nature of dark energy (e.g. Laureijs et al. 2011), and combined with redshift-space distortion measurements, test theories of gravity (e.g. Linder 2008).

From the perspective of galaxy formation, the clustering of galaxies informs us about the relation between galaxies and the underlying dark matter and can provide hints about the physical processes involved in galaxy assembly history. As galaxies reside within the dark matter haloes, their positions trace the underlying cosmic structure. While the formation and evolution of the dark matter haloes is governed exclusively by gravitational interaction, the assembly of the galaxies is governed by the more complex baryon physics that also affects the distribution of galaxies. Such ‘galaxy bias’ may impact as well cosmological inferences made from galaxy clustering measurements.

The main statistical tool used for characterizing galaxy clustering is the two-point correlation function \( \xi(r) \), which measures the excess probability over random of finding pairs of galaxies at different separations \( r \) (Peebles 1980). Commonly, when analysing redshift surveys, the projected correlation function integrated along the line of sight is used, in order to eliminate, in principle, redshift-space distortions (Davis & Peebles 1983). Observations show that brighter, redder and more massive galaxies are more strongly clustered and related trends are also measured as a function of morphology and spectral type (e.g. Norberg et al. 2002a; Zehavi et al. 2002, 2005; Goto et al. 2003; Coil et al. 2006; Li et al. 2006; Croton et al. 2007; Zheng, Coil & Zehavi 2007; Coil et al. 2008; Zehavi et al. 2011; Coupon et al. 2012; Guo et al. 2013; Farrow et al. 2015).

The theoretical modelling of galaxies plays an important role in interpreting clustering data, since, although galaxy clustering on large enough scales is very similar to that of the underlying matter distribution, it is not expected to be identical. Such models are also routinely used to estimate sample variance and verify methods...
for correcting observational biases. Several theoretical schemes are able to model galaxy clustering in volumes comparable to those probed observationally. These start from a dark matter only (DMO) $N$-body simulation, and populate haloes or sub-haloes with galaxies. Halo occupation distribution models (HOD; e.g. Cooray 2002; Berlind & Weinberg 2002; Tinker et al. 2012) or sub-halo abundance matching (SHAM, e.g. Vale & Ostriker 2004, 2006; Conroy, Wechsler & Kravtsov 2006) are statistical techniques that match galaxies with haloes by abundance based on, for example, their circular velocity. Semi-analytical galaxy formation techniques (e.g. Kauffmann, White & Guiderdoni 1993; Cole et al. 1994) use physically motivated schemes to associate galaxies with haloes (see e.g. Baugh 2006, for a review).

Notwithstanding the successes of these methods, they suffer from intrinsic limitations. Stellar and active galactic nucleus (AGN) feedback from forming galaxies affect the mass of their halo (Sawala et al. 2013; Velliscig et al. 2014; Schaller et al. 2015), limiting the extent to which any DMO simulation predicts the clustering of haloes as function of mass accurately. Feedback effects are plausibly strong enough to affect the mass distribution itself (Sembolini et al. 2011; van Daalen et al. 2011) and with it galaxy clustering (Hellwing et al. 2016). These effects may be relatively small, but the main limitation of the models is on smaller scales, where several effects that occur when a galaxy becomes a satellite (tidal interactions, ram-pressure stripping, strangulation, etc.) come into play. van Daalen et al. (2014) investigate how physics behind galaxy formation affect the clustering of galaxies at small scales through comparing two models of OWLS project (Schaye et al. 2010) with and without AGN feedback. They found that the physics of galaxy formation affects clustering on small scales, and, in addition, affects larger scales through its impact on the masses of sub haloes. Farrow et al. (2015) show how clustering in galaxy light cone mocks generated with two versions of the semi-analytical GALFORM code (Gonzalez-Perez et al. 2014; Lacey et al. 2016) differ significantly with the observations on small scales. McCullagh et al. (2017) and Gonzalez-Perez et al. (in preparation) show that once a more detailed merger scheme is considered, such as that described by Simha & Cole (2013), reasonably good agreement with clustering data on small scale is achieved. This is in agreement with the findings of Contreras et al. (2013) in which different families of galaxy formation models were compared to clustering data. McCarthy et al. (2017) present results from the BAHAMAS cosmological hydrodynamical simulation. These simulations were designed using a similar calibration strategy as EAGLE but using the OWLS implementation of galaxy formation described by Schaye et al. (2010). The simulated volume (400 Mpc $h^{-1}$ on a side) and mass resolution (initial baryonic particle mass $8 \times 10^4 M_\odot / h$) allow them to probe the galaxy correlation function on large scales, but not to go to lower mass galaxies and smaller scale clustering that we concentrate on here, or to investigate clustering as a function of galaxy property such as colour. Sales et al. (2015) compare the distribution of satellite galaxies from the ILLUSTRIS simulation (Vogelsberger et al. 2014) as a function of colour to Sloan Digital Sky Survey (SDSS; York et al. 2000) observations by Wang et al. (2014). They attribute the better agreement of the simulations compared to semi-analytical models to the more realistic gas contents of satellites at infall.

Galaxy clustering on small scales, and as a function of intrinsic galaxy properties such as luminosity, colour and star formation rate (SFR), thus might prove to be a stringent test of galaxy formation models. Performing such a test is the aim of this paper: We explore the clustering of galaxies in the cosmological hydrodynamical EAGLE simulation (Schaye et al. 2015) and its dependence on galaxy properties. The galaxy formation model of EAGLE uses sub-grid modules that are calibrated to reproduce the present-day stellar mass function (as described by Crain et al. 2015). In addition, the EAGLE simulation reproduces relatively well the colours and luminosities of galaxies both in the infrared (Camps et al. 2016) and at optical wavelengths (Trayford et al. 2016).

The 100-Mpc extent of the largest EAGLE simulation volume analysed in this paper is too small to properly sample large-scale modes, and, as is well known, such missing large-scale power quantitatively affects clustering measures even on smaller scales (e.g. Bagla & Ray 2005; Bagla & Prasad 2006; Trenti & Stiavelli 2008). To estimate the severity of this, we compare the clustering of haloes in a DMO version of the EAGLE volume to that in a much larger volume, simulated with the same cosmological parameters. This allows us to estimate the limitations of our approach. We compare the EAGLE predictions to clustering measurements by Farrow et al. (2015) of galaxies in the Galaxy and Mass Assembly redshift survey (GAMA; Driver et al. 2011; Liske et al. 2015), which are in accord with the Zehavi et al. (2011) SDSS measurements. For completeness, we note that Crain et al. (2017) show that EAGLE reproduces the observed clustering of $z = 0$ H i sources.

This paper is organized as follows: In Section 2, we describe the main characteristics of the simulations used, and briefly discuss the GAMA survey to which we compare. In Section 3, we define the notation and present the tools used to measure galaxy clustering. Simulations and observations are compared in Section 4. In the discussion, Section 5, we compare EAGLE with the clustering in GALFORM (following the analysis of Farrow et al. 2015) and our own clustering measurements using the data base of the ILLUSTRIS (Vogelsberger et al. 2014) simulation. The conclusions are summarized in Section 6.

Throughout this paper and unless specified otherwise, we use the Planck Collaboration XVI (2014) values of the cosmological parameters ($\Omega_b = 0.0482$, $\Omega_{dark} = 0.2588$, $\Omega_L = 0.693$ and $h = 0.6777$, where $H_0 = 100 h \, \text{km} \, \text{s}^{-1} \, \text{Mpc}^{-1}$). Observational measures of clustering are (most commonly) specified in $h^{-1}$-dependent units, and to ease the comparison to other clustering studies, we will express distances in $h^{-1}$ Mpc and masses in $h^{-1} M_\odot$.

## 2 Simulations and Data

This section briefly describes the simulations used and the GAMA survey to which clustering results of the simulations are compared.

### 2.1 The EAGLE hydrodynamical simulation suite

We use the ‘reference’ EAGLE simulation from table 2 in Schaye et al. (2015) (i.e. L0100N1504, but hereafter referred to as the EAGLE simulation), a hydrodynamical cosmological simulation that starts at $z = 127$ from initial conditions generated using the PANTHIA multi-resolution phases of Jenkins & Booth (2013), taking the Planck Collaboration XVI (2014) cosmological parameter values. The simulation is part of the EAGLE simulation suite (Crain et al. 2015; Schaye et al. 2015) and Table 1 lists some of the key simulation parameters. The EAGLE simulations were performed with the GADGET-3 code, which is based on GADGET-2 (Springel 2005), and uses ‘sub-grid models’, briefly discussed in more detail below, to encode physical processes below the resolution limit. These models are formulated using parameters or functional forms that express limitations in our understanding of a given process (for example, star formation) or our inability to simulate accurately a known process because of lack of numerical resolution (e.g. the effect of a
Table 1. Numerical parameters of cosmological simulations considered. From the left- to right-hand side: simulation identifier, simulation comoving side length \(L\), initial mass \(m_0\) of baryonic particles, dark matter particle mass, Plummer-equivalent comoving length \(\epsilon_{\text{com}}\) and maximum proper gravitational softening length \(\epsilon_{\text{prop}}\). The EAGLE and EAGLE-DMO simulations are referred to as L0100N1504 and L0100N1504-DMO by Schaye et al. (2015).

| Name        | \(L\) (\(h^{-1}\)Mpc) | \(m_0\) (10^5 M_\odot) | \(m_{\text{dm}}\) (10^7 M_\odot) | \(\epsilon_{\text{com}}\) (kpc) | \(\epsilon_{\text{prop}}\) (kpc) |
|-------------|------------------------|--------------------------|---------------------------------|-------------------------------|-------------------------------|
| EAGLE       | 67.77                  | 1.81                     | 9.70                            | 2.66                          | 0.7                           |
| EAGLE-DMO   | 67.77                  | –                        | 11.51                           | 2.66                          | 0.7                           |
| P-MILLENNIUM| 542.16                 | –                        | 157                             | 3.40                          | 3.40                          |

supernova explosion on the interstellar medium in the presence of radiative cooling). These parameters and functions are calibrated so that the simulation reproduces a limited set of observed properties of galaxies, by performing a large set of simulations in which these parameters are varied as described by Crain et al. (2015). The set of constraints is limited, mainly because each simulation takes a long time to run. In the case of EAGLE, sub-grid parameters were calibrated to observations at \(z \approx 0\) of the galaxy stellar mass function of Baldry et al. (2012), galaxy sizes as measured by Shen et al. (2003), and the relation between black hole mass and stellar mass.

The hydrodynamics used in EAGLE uses a number of improvements to the smoothed particle hydrodynamics implementation collectively referred to as ANARCHY and described by Dalla Vecchia (in preparation); see Schaller et al. (2015) for a discussion of the relatively small impact of these changes on the properties of simulated galaxies. We briefly summarize below the sub-grid modules for unresolved physics relevant for this paper:

(i) Photoheating and radiative cooling by the optically thin evolving ultraviolet UV/X-ray background of Haardt & Madau (2001) is implemented element-by-element as described by Wiersma, Schaye & Smith (2009a).

(ii) Star formation is implemented using the pressure law of Schaye & Dalla Vecchia (2008) and the metallicity-dependent star formation threshold of Schaye (2004). Gas particles eligible for star formation are converted to ‘star particles’ stochastically with a probability that depends on their SFR and their time-step.

(iii) Stellar evolution and enrichment is implemented as described by Wiersma et al. (2009b): Assuming that stars form with the Chabrier (2003) stellar initial mass function (IMF), spanning the range [0.1,100] M_\odot, we use stellar evolution and yield tables to calculate the rate of type Ia and type II (core-collapse) supernovae, and follow the rate at which stars enrich the interstellar medium through asymptotic giant branch, type Ia and type II evolutionary channels.

(iv) Seeding, accretion and merging of black holes is implemented following Springel, Di Matteo & Hernquist (2005) and Booth & Schaye (2009), modified to account for the angular momentum of accreted gas as described by Rosas-Guevara et al. (2015).

(v) Thermal feedback from stars is implemented as described by Dalla Vecchia & Schaye (2012); feedback from accreting black holes is also implemented thermally.

(vi) Dark matter haloes are identified using the friends-of-friends (FoF) algorithm, with baryonic particles (gas, stars and black holes) assigned to the same halo as the nearest dark matter particle, if any. The mass of the halo is characterized by \(M_h \equiv M_{200,c}\), the mass enclosed within a sphere within which the mean density is 200 times the critical density.
population using
\[(g - r)_b = 0.7 - 0.028(M_{\text{vir}} + 18.6) \quad (\text{EAGLE red-blue cut}), \tag{2}\]
shown as the black line in Fig. 1. Therefore, the aim of the colour cut is to separate the passive from the active population of galaxies. It is essential to ensure that comparable cuts are made in the different data sets. The slopes of observed and simulated colour cuts are virtually identical, and they are offset in colour by less than 0.1 mag at \(M_{\text{vir}} = -18.6\). This offset is comparable to the offset of \(-0.15\) (but in the opposite direction to the EAGLE offset) for the semi-analytical model considered by Farrow et al. (2015) (see their fig. 2).

2.2 DMO simulations

We use two DMO simulations to study the impact of the limited simulation volume of the EAGLE simulation on clustering: one with the same volume and initial conditions as EAGLE (referred to as EAGLE-DMO below, this is the simulation L00100N1504-DMO described by Schaye et al. 2015) and one with a much larger simulation volume (referred to as P-MILLENNIUM below). Combined, they allow us to assess to what extent missing large-scale power and sample variance affect inferences on clustering from the relatively small EAGLE volume.

Simulation EAGLE-DMO has the same volume, gravitational softening, cosmology and initial conditions as EAGLE. The masses of the dark matter particles are increased by a factor of \((\Omega_b + \Omega_{\text{dark}})/\Omega_{\text{dark}}\) compared to EAGLE, to account for not including the baryonic mass. We use this simulation to study clustering of haloes compared to other models – without results being affected by galaxy formation. The impact of baryonic effects on the density profiles of haloes in EAGLE was investigated by Schaller et al. (2015).

The P-MILLENNIUM simulation (see Table 1; Baugh et al., in preparation, and McCullagh et al. 2017) uses identical cosmological parameters as EAGLE but has a much larger volume (800 \(^3\) Mpc\(^3\) for the EAGLE simulation used here). With P-MILLENNIUM, we can quantify the effects of missing large-scale power and poor sampling of long wavelengths on clustering statistics by comparing the dark matter halo clustering in EAGLE-DMO with P-MILLENNIUM. For details about P-MILLENNIUM, see Baugh et al., in preparation, and McCullagh et al. (2017).

In both EAGLE-DMO and P-MILLENNIUM, dark matter haloes were identified using the FoF algorithm with the standard value of \(b = 0.2\) for the linking length in units of the mean particle separation. The mass of the halo is represented by \(M_{\text{vir}}\), the mass enclosed within a sphere with a density 200 times the critical density.

2.3 The GAMA survey

To put the EAGLE simulation results into context, we will compare them primarily to results from the GAMA survey data, and in particular to clustering measurements made by Farrow et al. (2015). The GAMA survey (Driver et al. 2011; Liske et al. 2015) is a spectroscopic and multiband surveys of galaxies carried out on the Anglo-Australian telescope. This work, we make use of the main \(r\)-band limited data from the GAMA equatorial regions \((\sim 180 \text{ deg}^2)\), which consists of a highly complete \((\sim 98 \text{ per cent})\) spectroscopic catalogue of galaxies selected from the SDSS DR7 (Abazajian et al. 2009) to \(r_{\text{petrosian}} < 19.8\). Further details of the GAMA survey input catalogue, tiling algorithm, redshifting and survey progress are described by Baldry et al. (2010), Robotham et al. (2010), Baldry et al. (2014) and Liske et al. (2015), respectively.

For the clustering comparisons presented in Section 4, we are primarily interested in the following GAMA galaxy properties: \(r\)-band absolute magnitude, stellar mass and rest-frame \((g - r)\) colour. We describe in turn how each of those properties have been estimated in the clustering measurements of Farrow et al. (2015):

(i) \(r\)-band absolute magnitude \((M_r)\): We apply evolution corrections and \(k\)-corrections to \(z_{\text{crit}} = 0\) Petrosian absolute magnitudes, where \(M_r \equiv M_r - 5 \log_{10}(h)\). For further details, see section 2.1.4 of Farrow et al. (2015).

(ii) \((g - r)\) colour: The rest-frame colours are derived from SDSS model magnitudes, with colour- and redshift-dependent \(k\)-corrections as per Loveday et al. (2012) and McNamara-Roberts et al. (2014). For further details, see section 2.1.4 of Farrow et al. (2015).

(iii) Stellar mass \((M_*)\) in units of \(h^{-2} M_\odot\): The clustering measurements in Farrow et al. (2015) use the relation between rest-frame \((g - i)\) colour and stellar mass as derived by Taylor et al. (2011), using the Bruzual & Charlot (2003) synthetic stellar population models with a Calzetti et al. (2000) dust attenuation law to correct for dust in the Milky Way. For further details, see section 2.1.5 of Farrow et al. (2015) and Taylor et al. (2011).

For completeness, we refer the reader to section 3 of Farrow et al. (2015) for the modelling of the GAMA selection function. Accurate modelling of the selection function of a redshift survey is key for precise clustering measurements. The modelling approach described in Farrow et al. (2015), which is based on the method of Cole (2011), enables a uniform modelling for all galaxy samples split by stellar mass, luminosity and colour.

3 GALAXY CLUSTERING

In this section, we present the analysis methods used, starting with the estimators we use for calculating the two-point correlation function and its associate errors. We then briefly discuss how we compute the effective bias.

3.1 The two-point correlation function

The spherically averaged two-point correlation function, \(\xi(r)\), defined as (e.g. Peebles 1980)
\[\xi(r) = \frac{1}{\langle n \rangle} \frac{dP}{dV} - 1,\tag{3}\]
provides a statistical description of a sample's spatial distribution. Here, \(dP/dV\) is the probability of finding a galaxy in the volume \(dV\) at a given (comoving) distance \(r\) from another galaxy, and \(\langle n \rangle\) is the mean (comoving) number density of galaxies. In practice, for a volume with periodic boundary conditions, the correlation function can be estimated by counting the number of pairs of galaxies, \(N_2(r)\), in a shell of volume, \(V_s(r)\), at distance \(r\) from each other using (e.g. Rijvo 1986)
\[\xi(r) = \frac{1}{\langle n \rangle^2 V_s(r)} N_2(r) - 1,\tag{4}\]
where \(V\) is the total volume of the periodic simulation.

However, when the volume has boundaries (as is the case for any observed survey, or when we want to restrict the analysis to a small region within a larger periodic simulation volume), these equations cannot be used. In such cases, \(\xi\) can be computed by comparing the distribution of galaxy pairs to the clustering of a set of points.
uniformly distributed within the survey volume, using, for example, the Landy & Szalay (1993) estimator:

\[ \xi(r) = \frac{DD - 2DR + RR}{RR}, \tag{5} \]

where DD is the suitably normalized number of galaxy pairs at a distance \( r \) from each other, RR is the corresponding normalized number of pairs from the random distribution and DR is the suitably normalized numbers of galaxies and random pairs separated by distance \( r \).

The comoving distance between galaxies cannot be measured directly from a redshift survey due to galaxy-peculiar velocities and large-scale redshift-space distortions. However, by splitting the information into a projected separation, \( r_p \), and distance parallel to the line of sight, \( \pi \), one can estimate the two-dimensional (2D) correlation function, \( \xi(r_p, \pi) \), which, in turn, is used to estimate the projected correlation function

\[ w_p(r_p) = 2 \int_0^{\pi_{\text{max}}} \xi(r_p, \pi) \, d\pi, \tag{6} \]

with \( \pi_{\text{max}} \) set to a value adequate for the sample considered (here \( \pi_{\text{max}} \) is fixed to \( \sim 34 \, \text{Mpc} \, h^{-1} \), which represents \( \sim L/2 \) of EAGLE simulation; see Table 1). We select \( \pi_{\text{max}} \) to be sufficiently large to account for most redshift-space distortions. In addition, the \( \pi_{\text{max}} \) value chosen is in line with what is commonly used in observational clustering measurements. To a very good approximation, \( w_p(r_p) \) is independent of redshift-space effects, making this statistic ideal for model comparisons. Furthermore, we tested the systematic differences between \( \xi(r) \) and \( w_p(r_p) \) and their dependence on \( \pi_{\text{max}} \), finding that the systematic difference is significantly smaller than the statistical errors. We compute \( w_p(r_p) \) along three orthogonal directions in the simulations, to improve the signal-to-noise ratio of the clustering measurements.

To reduce the dynamical range when plotting \( w_p \), we will often divide it by the projected correlation function of the reference power law, \( \xi(r) = (r_0/r)^{\gamma} \), with \( r_0 = 5.33 \, h^{-1} \, \text{Mpc} \) and \( \gamma = 1.8 \), from Zehavi et al. (2011) for the galaxy sample with \( -21.0 < M_{h,0} < -20.0 \), and where the constants are from the fitted function that corresponds to this power law as

\[ w_p^{\text{ref}}(r_p) = r_p \left( \frac{r_0}{r_p} \right)^\gamma \frac{\Gamma(1/2)\Gamma(\gamma - 1/2)}{\Gamma(\gamma / 2)}, \tag{7} \]

where \( \Gamma \) denotes the Gamma function.

### 3.2 Error estimates on clustering statistics

We compute and quote jackknife errors on the simulated two-point correlation function in EAGLE to mimic observational errors. However, sample variance is likely to dominate the error budget. We estimate sample variance by sub-sampling EAGLE-sized-volumes in the P-MILLennium simulation, which also allows us to examine any effects due to missing large-scale power. Unfortunately, we can estimate these errors for only the clustering of haloes – not galaxies – since the P-MILLennium simulation is DMO.

We apply the jackknife technique by partitioning the EAGLE (or EAGLE-DMO) simulation volume in \( N_{\text{sub}} \) tiles of equal volume, with \( N_{\text{sub}} = 8 \). We then compute the two-point correlation function \( \xi^{\text{JK}}(r) \) by omitting the \( i \)-th tile, and compute the variance

\[ \sigma^2(r) = \frac{(N_{\text{sub}} - 1)}{N_{\text{sub}}} \sum_{i=1}^{N_{\text{sub}}} \left( \xi^{\text{JK}}(r) - \xi^{\text{MK}}(r) \right)^2, \tag{8} \]

where \( \xi^{\text{MK}}(r) \) is the correlation function of the total volume. Such jackknife error estimates have been used extensively to estimate errors in galaxy clustering (e.g. Zehavi et al. 2002, 2011; Favole et al. 2016), and Zehavi et al. (2002) shows that such errors accurately reflect uncertainties in the clustering on the scales investigated here. However, it has, of course, its limitations, as pointed out by, for example, Norberg et al. (2009). For example, the technique may underestimate errors when a few systems dominate the signal. We will see below that this is in fact the case in EAGLE, where the clustering of low-mass red galaxies on small scales is dominated by satellites in a few massive clusters, as we illustrate in Fig. 7 below.

We estimate errors on the clustering of haloes due to sample variance and missing large-scale power on EAGLE volumes using the P-MILLennium simulation as follows: We partition the P-MILLennium simulation in \( N_{\text{sub}} = 512 \) tiles of volume equal to that of EAGLE. We then calculate the correlation function of haloes in the \( i \)-th tile, \( \xi_i(r) \), using equation (5), as well as the correlation function of the total volume, \( \xi(r) \). The variance is then calculated as

\[ \sigma^2(r) = \frac{1}{N_{\text{sub}} - 1} \sum_{i=0}^{N_{\text{sub}}} (\xi_i(r) - \xi(r))^2. \tag{9} \]

We use the number density of haloes of each non-overlapping tile to compute \( \xi(r) \).

### 3.3 Effective bias

The bias, \( b \), of a tracer population is the ratio of the correlation function of that tracer over that of the mass (e.g. Davis et al. 1985), with the effective bias \( b_{\text{eff}}(X) \), given by (e.g. Porciani, Magliocchetti & Norberg 2004)

\[ b_{\text{eff}}(X) = \frac{\int b(M_h) N_{\text{gal}}(M_h, X) n(M_h) \, dM_h}{\int N_{\text{gal}}(M_h, X) n(M_h) \, dM_h}. \tag{10} \]

where \( n(M_h) \) is the halo mass function (the number density of haloes of mass \( M_h \)), \( b(M_h) \) is the linear bias factor of haloes of mass \( M_h \) and \( N_{\text{gal}}(M_h, X) \) is the mean number of galaxies of property \( X \) in haloes of mass \( M_h \) (the mean halo occupation). The property \( X \) could select galaxies in a given stellar mass, colour or luminosity range, for example. We approximate the integral as a sum over all haloes in the simulation, obtaining

\[ b_{\text{eff}}(X) = \frac{\sum_{j=1}^{N_{\text{halo}}} b_j(M_j) N_{\text{gal}}(M_j, X)}{\sum_{j=1}^{N_{\text{halo}}} N_{\text{gal}}(M_j, X)}, \tag{11} \]

where \( N_{\text{gal}}(M_j, X) \) is the number of galaxies with property \( X \) in a halo of mass \( M_j \). In practice, we estimate the effective bias of samples split by stellar mass and hence evaluate this sum for all galaxies in narrow stellar mass bins.

To estimate the linear halo bias \( b(M_h) \), we follow Mo & White (1996):

\[ b(M_h) = 1 + \left( \frac{\nu(M_h)^2}{\sigma(M_h)} - 1 \right)/\delta_c, \tag{12} \]

where \( \delta_c = 1.686 \) is the spherical collapse density threshold and \( \nu(M_h) = \delta_c/\sigma(M_h) \) is the dimensionless amplitude of fluctuations that produce haloes of mass \( M_h \) (at a given redshift \( z \)). The (linear) matter variance, \( \sigma(M_h) \) at a given \( z \), can be computed numerically for a given linear power spectrum (see equation 9 in Murray, Power & Robotham 2013) using the web-portal HMFcalc 1 (and adopting the spectral index \( n_s = 0.9611 \) used in EAGLE). Finally, we adopt the fit

1 http://hmf.icrar.org
provided by Jenkins et al. (2001) for the halo mass function, \(n(M_h)\), which provides a good description of the halo mass function of all simulations used here.

4 RESULTS

We begin by considering the clustering of dark matter haloes in EAGLE, followed by a quick look at the real- and redshift-space clustering of EAGLE galaxies with well-resolved stellar mass. In Section 4.3, we present the main results of this study, namely the clustering of galaxies in EAGLE compared to that of GAMA, when split by stellar mass, luminosity, colour or SFR.

4.1 Halo clustering in EAGLE

The real-space clustering of haloes with \(M_h > 10^{12} h^{-1} M_\odot\) in the P-MILLENIUM dark matter simulation is plotted in Fig. 2. We subdivide the volume of this simulation into 512 non-overlapping tiles, each with the same volume as EAGLE-DMO, and compute the correlation function \(\xi(r)\) for each of the tiles, using the mean number density of haloes in each tile in equation (5). We plot the mean correlation function averaged over all tiles, \(\bar{\xi}(r)\), as the grey line, with the scatter around the mean shown as 1\(\sigma\) error bars. The mean correlation function \(\xi(\bar{r})\) follows the correlation function of the full volume (black dashed line) very closely, falling well within the scatter between volumes, as expected.

The correlation function of EAGLE-DMO (green line) falls below \(\bar{\xi}(r)\) on scales smaller than 1\(h^{-1}\) Mpc, remains well within sample variance up to scales 5–6\(h^{-1}\) Mpc (log \(r/\bar{h}^{-1}\) Mpc = 0.7–0.78), then falls increasingly below \(\bar{\xi}(r)\) above this scale. As both simulations have identical power spectra and cosmological parameters, the deviations are due to sample variance and due to the integral constraint on \(\xi\). We note that numerical resolution is not likely to play a role in the apparent differences in clustering, as these haloes are resolved by \(\approx 10^4\) particles or more.

We compute jackknife errors for EAGLE-DMO, as described above, and plot them in green. Of course these quantify neither finite-volume effects nor sample variance, but nevertheless, the green and black curves are within the EAGLE-DMO jackknife errors. This motivates us to use such errors when calculating errors on the correlation function of galaxies, rather than haloes, below, since we do not have the access to larger hydrodynamical simulations to estimate finite-volume effects. Given the level of convergence between EAGLE-DMO and P-MILLENIUM, we will plot correlation functions up to scales of 10\(h^{-1}\) Mpc (14 per cent of the full extent), with Fig. 2 quantifying the limitations on halo clustering.

4.2 Galaxy clustering in EAGLE

The real-space correlation function, \(\xi(r)\), of galaxies with stellar mass \(M_\ast > 10^{8.83} h^{-2} M_\odot\) from EAGLE is plotted in Fig. 3 (top panel). Distinguishing between central galaxies (typically but not necessarily the most massive galaxy in a given halo) and satellites galaxies, we compute the one- and two-halo contributions separately (grey dashed and grey dotted lines, respectively). The red line is the reference power-law model for galaxies with \(-21.0 < M_\ast < -20.0\) from the fit by Zehavi et al. (2011). Bottom panel: corresponding projected correlation function, \(w_p(r_p)\) from equation (6). The red line is the projected correlation function from the fit by Zehavi et al. (2011), \(w_p^{\text{ref}}\) from equation (7). The galaxy selection is different in detail for EAGLE and the observations: The red lines are shown to guide the eye. [A colour version of this figure is available in the online version.]
this does not imply that the same is true for the galaxy correlation function, since there are, of course, generally many galaxies per halo.

The real-space correlation function quantifies the physical clustering of galaxies, independent of any peculiar velocities. However, observations can only measure clustering only in redshift space, and peculiar velocities then distort the signal. To ameliorate the effects of such redshift-space distortion, it is convenient to integrate \( \xi(r_p, \pi) \) over a narrow range in \( \pi \), and compute the projected correlation function \( w_p(r_p) \); see equation (6). This is plotted for the same EAGLE galaxies (those with \( M_*>10^{8.83}\ h^{-2}\ M_\odot \)) in the bottom panel of Fig. 3, again plotting one- and two-halo terms as well.

The reference model \( w_p^{\text{ref}}(r_p) \) of equation (7) provides a relatively good fit to EAGLE’s projected correlation function. We note, however, that the galaxy selections differ between EAGLE and the SDSS galaxies fitted by Zehavi et al. (2011), which gives rise to \( w_p^{\text{ref}}-\text{two} \) the two therefore did not have to agree: We show the comparison to guide the eye and because we use \( w_p^{\text{ref}} \) as a normalization below.

The contributions of one- and two-halo terms to \( w_p \) are equal at a projected separation of \( r_p \sim 0.4\ h^{-1}\ Mpc \). Comparing the top and bottom panels from Fig. 3, it is clear that the two-halo term contributes more to the projected correlation function on scales comparable to the virial radius of haloes than it does to the two-point correlation function: At a scale of \( \approx0.1\ h^{-1}\ Mpc \), the two-halo term contributes nearly 10 per cent to \( w_p \).

The 2D redshift-space correlation function of this stellar-mass-limited sample of EAGLE galaxies is plotted in Fig. 4 in terms of the projected separation, \( r_p \), and line-of-sight separation, \( \pi \). It exhibits the familiar elongation in the \( \pi \) direction at small \( r_p \) resulting from virial motion of galaxies in haloes (the ‘fingers-of-god’ effect), and the flattening in the \( \pi \) direction at large \( r_p \), due to coherent streaming motions of galaxies into haloes and out of voids (the ‘Kaiser’ effect; Kaiser 1987). We do not compare this correlation function directly to GAMA, mainly because of the complexity of making sure that the selection of galaxies in the \( \pi \) direction is the same in simulation and data.

4.3 Galaxy clustering in EAGLE compared to GAMA

We use the volume-limited samples of GAMA galaxies presented by Farrow et al. (2015), which we can split by stellar mass, luminosity or colour. For samples split in bins of stellar mass or luminosity only, we refer the reader to Table 2 of Farrow et al. (2015), while we present in Table 2 the properties of the additional GAMA samples used in here, for which we have computed clustering statistics following the methods outlined by Farrow et al. (2015).

Throughout this section, EAGLE galaxies are selected from the \( z=0.1 \) snapshot (a redshift close to the median redshift of the GAMA samples). Some statistics of the EAGLE samples used are provided in Table 3. By construction and as explained in Section 2.1, the galaxy formation model in EAGLE yields a stellar mass function that is in relatively good agreement with that inferred from GAMA in the mass range we analyse here. The agreement is not as good as in statistical methods that populate dark matter haloes with galaxies, such as SHAM or HOD, for example, and which yield the correct number densities by construction. Interestingly, however, EAGLE predicts a scatter in stellar mass at a given halo mass that depends on halo concentration, although the dependence is not strong enough to explain the full variance (Matthee et al. 2017). Such non-linear dependences are not taken into account in these statistical methods.

Hence, because the mean number density of galaxies in each bin of stellar mass, agree reasonably well between data and simulation, we do not compare clustering at given number density, as commonly done, but directly compare samples selected by stellar mass - or indeed luminosity.

4.3.1 Stellar-mass-dependent clustering

A comparison of the clustering of EAGLE galaxies to that in GAMA as a function of stellar mass is shown in Fig. 5, using the same mass as used by Farrow et al. (2015). The bottom panels of Fig. 5 highlight the differences between the two measurements, by presenting the ratio of the projected correlation functions with respect to the reference power-law model adopted (following Farrow et al. 2015). The projected correlation function of EAGLE galaxies (green lines with jackknife errors) is in remarkably good agreement with the GAMA data (solid black lines, with 1σ uncertainty range shown as a grey shaded area): The deviations are typically within the measured uncertainty range.

It is well known (and clear from the figure) that the clustering strength of galaxies increases with stellar mass. There is little or no evidence for such a trend in the simulation; however, the errors are relatively large and the simulation is not inconsistent with such a trend either. Furthermore, the size of the simulation prevents us to test the stellar-mass-dependent clustering, due to observed trends that are visible only when having a large dynamic range of stellar masses. Therefore, a larger volume would be needed to confirm such trend. The good agreement in shape of the correlation functions of GAMA and EAGLE is encouraging.

The calibration of sub-grid parameters in EAGLE was based on one-point statistics, as described in Section 2. The clustering of galaxies is therefore a genuine model prediction. The good agreement then implies that EAGLE galaxies tend to inhabit haloes in a way that

\[ \text{Farrow et al. (2015) use a flat } \Omega_m = 0.25 \text{ cosmology to infer distances for their clustering measurements. On the scales considered, this difference in cosmology is totally negligible.} \]
mimics accurately the way GAMA galaxies do. Finally, we note that the decrease of the clustering signal on scales greater than 5 h⁻¹ Mpc scales in EAGLE is related to the limited simulation box size – and mimics the corresponding fall in clustering of EAGLE haloes.

4.3.2 Luminosity-dependent clustering

A comparison of the clustering of EAGLE galaxies to that in GAMA as a function of r-band luminosity is shown in Fig. 6. The agreement is very good, and well within the relatively large jackknife error estimates. Similar to the case of clustering as a function of mass, the amplitude of observed clustering increases with luminosity (see e.g. Norberg et al. 2001; Zehavi et al. 2005), but again there is little or no evidence for such a trend in EAGLE. As in the previous section, we suggest this is mostly due to the finite volume of the simulation: More luminous galaxies are biased to more massive haloes, of which there are relatively few in the EAGLE volume, and their clustering is underestimated because of lack of large-scale power. The reference power law \( w_p^{ref} \) from equation (7) describes the clustering of EAGLE

**Table 2.** Statistics of GAMA galaxy samples (mostly volume-limited) that are not already described in table 2 of Farrow et al. (2015), following a similar table structure. The stellar-mass-restricted samples are split into red and blue galaxies, presented in turn. Columns from the left to right-hand side are as follows: stellar mass range, number of galaxies, galaxy number density (including JN errors), fraction of satellites, fraction of blue galaxies (following equation 2), the SFR limits as a function of redshift, total number of galaxies, the galaxy number density, median sample redshift, median r-band absolute magnitude, median stellar mass, median (g - r)₀, rest-frame colour and the fraction of truly volume-limited galaxies (see the text for further details).

**Table 3.** Statistics of EAGLE stellar-mass-selected samples from the \( z = 0.1 \) snapshot. Columns from the left- to right-hand side are as follows: stellar mass range, number of galaxies, galaxy number density (including JN errors), fraction of satellites, fraction of blue galaxies (following equation 2), the SFR limits as a function of stellar mass; the mass bin is indicated in each column. Top panels: Projected correlation function in red and blue galaxies. Bottom panels: same as top panels, but the correlation functions are divided by the reference function \( \bar{w}(r_p) \) from equation (6).
consider this sample still to be volume-limited when computing clustering statistics. The advantage of keeping the exact same volumes for all stellar mass samples split by colour (hence, they all sample the same underlying large-scale structure) overcomes this minor subtlety, which is primarily driven by uncertainties in the measured colours and adopted k-corrections for a small subset of galaxies.

We computed colours of EAGLE galaxies as explained in Section 2.1. The colour–magnitude diagram exhibits a blue cloud of star-forming galaxies, well separated from a ‘red sequence’ of passive galaxies, as shown by Trayford et al. (2015) (see also Trayford et al. 2017). The colour of EAGLE’s red sequence is slightly bluer than in GAMA, which may due to differences in metallicity and/or limitations in the adopted population synthesis models, as discussed by Trayford et al. (2015). When comparing to GAMA, we want to study whether the clustering of star-forming (blue) galaxies differs from that of passive (red) galaxies, at a given mass. We therefore divide EAGLE galaxies in bins of stellar mass and colour using the same mass bins as used in the analysis of GAMA, but apply the slightly different colour cut to distinguish red from blue using equation (2), as compared to the cut of equation (1) applied to GAMA. Some statistics of the EAGLE galaxies are summarized in Table 3, including the fraction of EAGLE galaxies that are satellites.

The projected correlation functions for red and blue galaxies, split in bins of stellar mass, are plotted in Fig. 7. To ease the interpretation of the EAGLE clustering results, we show the correlation function for all central and satellite galaxies (i.e. irrespective of colour) within each stellar mass bin as green dotted and dashed lines, respectively. Therefore, the central/satellite clustering results are the same in the left- and right-hand panels, but vary with stellar mass (from the top to bottom).

In EAGLE, red galaxies (left-hand column) cluster more strongly than blue galaxies (right-hand panel) of a given mass, and similarly satellite galaxies (dashed lines) cluster more strongly than centrals (dotted lines), in all stellar mass ranges studied. It is also apparent that the red population follows closely the clustering of satellites, in particular for galaxies with stellar masses greater than $10^{10} \, h^{-2} \, M_\odot$. In contrast, blue galaxies follow more closely the clustering of centrals, again particularly for the two more massive galaxy stellar bins.

The clustering of blue EAGLE galaxies tracks that of blue GAMA galaxies well, in particular, on scales up to $r_p \sim 4 \, h^{-1} \, \text{Mpc}$ [$\log_{10}(r_p) = 0.6$. However, red EAGLE galaxies cluster noticeably more strongly than red GAMA galaxies. As a consequence, the trend that red galaxies cluster more strongly than blue galaxies of given mass, which is clearly present in GAMA, is too strong in EAGLE. Trayford et al. (2015) noticed that red galaxies are over-abundant in EAGLE at low mass, and demonstrated that this is at least partly due to lack of numerical resolution (see the appendix of Trayford et al. 2015). At higher stellar masses, EAGLE yields too large fraction of blue galaxies instead, plausibly a consequence of the dust screen not suppressing blue light from star-forming regions significantly (Trayford et al. 2017). We suspect therefore that it is the overly strong suppression of star formation in small galaxies as they become satellites, most likely as a consequence of lack of numerical resolution, which causes EAGLE to overpredict small-scale clustering of red galaxies. Consistent with this interpretation, we find that the strong clustering of red galaxies in EAGLE is significantly influenced by the presence of a few massive haloes. To demonstrate this, we re-compute $w_p$ for red EAGLE galaxies after excluding all galaxies in the three most massive haloes. We show the result by the orange line in Fig. 7. The overall amplitude of the clustering signal of red galaxies is dramatically reduced. A larger simulation box is likely

---

**Figure 6.** Same as the bottom panels of Fig. 5, but for clustering as a function of $r$-band luminosity. Ranges of $r$-band absolute magnitudes are shown from faint to bright (top to bottom panels), with the number of EAGLE galaxies in this luminosity range labelled in each panel. The jackknife error bars for EAGLE galaxies are included.
required to provide detailed insight on the clustering of red massive galaxies.

The colour dependences of the clustering of EAGLE galaxies of a given stellar mass is clearly partly due to the relative fractions of satellite and central galaxies. In Table 3, we show the fraction of satellites and blue galaxies for each stellar mass range. We find that the fraction of satellite galaxies decreases at higher masses, while the fraction of red and blue satellites is not strongly dependent on stellar mass. For example, in the stellar mass range $9.5 < \log M_*/M_\odot h^{-2} < 10.0$, we find that $\sim 19$ per cent are red and the remaining $\sim 81$ per cent are blue, while 43 per cent of the galaxies are satellites. The table also shows that the vast majority of galaxies are blue, but also that most of galaxies of the complete sample in this stellar mass bin are centrals. Furthermore, the good agreement with GAMA in the clustering of blue galaxies is consistent with the fact that EAGLE galaxies of given mass cluster similarly to GAMA galaxies, as shown in Fig. 5. A relatively modest improvement of the numerical resolution of the simulation could be enough to reproduce the clustering of red galaxies equally well (see, Trayford et al. 2015, for further details). However, the difference seen could equally be due to the missing large-scale power and the impact the few rare objects have in the 100-Mpc EAGLE volume.

### 4.3.4 SFR-dependent clustering

We divide EAGLE galaxies of a given stellar mass in three bins of SFR: the 30 per cent with the lowest star formation rate (‘low SFR’), the 30 per cent with the highest star formation rate (‘high SFR’) and the remainder (‘int SFR’). Their galaxy clustering is plotted in Fig. 8. The low SFR galaxies are clustered most strongly, which is particularly evident on the smaller scales, and this is the case in all stellar mass bins investigated. The difference in the amplitude of clustering between the high and intermediate SFR galaxies is not very large, with high SFR galaxies generally clustering least.
GAMA SFR measurements to be directly compatible with those in samples defined by absolute cuts in SFR. This, in turn, requires the tailed comparison to take place, it would be necessary to work with Gunawardahana et al. in preparation for details). Hence, for a de-

limitations in each of the GAMA and EAGLE star-forming samples ± typical colour version of this figure is available in the online version.]

Table 3 lists the corresponding cuts in SFR. [A (green line), the 30 per cent least star-forming galaxies ('low SFR', purple), panel. In each galaxy stellar mass bin, we show the result for all galaxies function (7), is plotted for galaxies in bins of stellar mass, as labelled in each 

Figure 8. Clustering as function of SFR. The projected galaxy correlation function of EAGLE galaxies, divided by the reference power law of equation (7), is plotted for galaxies in bins of stellar mass, as labelled in each panel. In each galaxy stellar mass bin, we show the result for all galaxies (green line), the 30 per cent least star-forming galaxies (‘low SFR’, purple), the 30 per cent highest star-forming galaxies (‘high SFR’, blue) and the remainder (‘int SFR’, red). Table 3 lists the corresponding cuts in SFR. [A colour version of this figure is available in the online version.]

Jackknife errors bars are not plotted to avoid clutter, but should be of the order of ±0.35 dex (i.e. a factor of \sqrt{3} higher than the typical ±0.2 dex errors seen in Fig. 5 below \( r_p = 1 \text{ h}^{-1} \text{Mpc} \)). Any difference in clustering between the blue (‘high SFR’) and red (‘int SFR’) curves is therefore not very significant. At the largest scales shown, the amplitude of clustering for all populations becomes similar, with any difference now much smaller than jackknife errors.

We do not compare these clustering measurements to those from GAMA presented by Gunawardhana et al. (in preparation), for the following main reason: The SFR and stellar mass range probed by GAMA and EAGLE are, in detail, poorly matched, in the sense that the limitations in each of the GAMA and EAGLE star-forming samples are hard to account for all at the same time. For GAMA, the SFR is only measurable for galaxies with sufficiently high SFR, resulting in volume-limited samples with SFR and stellar mass ranges restricted to SFR \( \gtrsim 0.3 \text{ M}_\odot \text{ yr}^{-1} \) and \( \text{M}_* \gtrsim 10^{9.5} \text{ h}^{-2} \text{ M}_\odot \) (see Gunawardahana et al. in preparation for details). Hence, for a detailed comparison to take place, it would be necessary to work with samples defined by absolute cuts in SFR. This, in turn, requires the GAMA SFR measurements to be directly compatible with those in EAGLE, as defining samples by SFR ranking is not possible. Although the galaxy number densities of GAMA and EAGLE split by stellar mass are in reasonably good agreement, splitting these sub-samples by SFR does not necessarily result in samples with similar number densities, due to differences in the bivariate SFR–M\* distribution. In fact, we find that the galaxy number densities in EAGLE are between 60 and 70 per cent larger than in GAMA for the same stellar mass and SFR range. This implies that a detailed clustering comparison becomes futile: Any difference observed in the clustering could be attributed to the differences in the measured number densities of the samples. This is in agreement with the results of Furlong et al. (2015), who pointed out that the specific SFRs in EAGLE are typically 0.2–0.5 dex lower than observed. A proper understanding of those SFR differences between data and simulations is required before a detailed and informative clustering comparison of SFR-selected samples can be made.

The differences in clustering of galaxies of the same mass that are satellites versus centrals (Fig. 7), or red versus blue galaxies (Fig. 8), should be dependent on the mass of the halo they inhabit, with more strongly clustered galaxies residing in more massive haloes. To verify that this is the case in EAGLE, we plot the HOD – the fraction of galaxies of a given \( M_\star \) that inhabit haloes of mass \( M_h \) – for galaxies in the stellar mass range \( 10^{-0.5} < M_\star / h^2 \text{ M}_\odot < 10^{10} \) split in centrals and satellites (Fig. 9), or in bins of SFRs for three ranges in \( M_\star \) (Fig. 10).

Central galaxies in the mass range \( 10^{0.5} < M_\star / h^2 \text{ M}_\odot < 10^{10} \) inhabit haloes with a narrow range of masses, from \( 10^{11.2} \) to \( \sim 10^{11} \text{ h}^{-1} \text{ M}_\odot \). In contrast, satellites of the same stellar mass inhabit haloes with a wide range of masses, from \( 10^{11} \text{ h}^{-1} \) to \( \sim 10^{14.5} \text{ h}^{-1} \text{ M}_\odot \) (Fig. 9). The (much) stronger clustering of satellites is therefore clearly due to the significant fraction that resides in these much more massive (and hence more clustered) haloes (see e.g. Guo et al. 2014).

At a given stellar mass, EAGLE galaxies with a higher SFR inhabit lower mass haloes than those with a lower value of SFR (at \( z = 0.1 \)), as shown in Fig. 10 for three ranges in \( M_\star \). The figure also demonstrates that the halo occupation is similar for galaxies with a high or intermediate SFR. This explains the results from Fig. 8 that, at given \( M_\star \), galaxies with a low SFR cluster more strongly than those with higher SFR.
The SFR (and hence also colour) dependence of clustering in EAGLE is related to the mechanism that causes some galaxies to have a low SFR for their mass; the reduction in SFR once a galaxy becomes a satellite, which is discussed in detail for the EAGLE simulations by Trayford et al. (2016). The reduction of the SFR of satellites results from two related physical processes that operate in hydrodynamical simulations: ram-pressure stripping, mainly of the outer parts of satellites as shown by Bahé & McCarthy (2015) using the GIMIC simulations (Crain et al. 2009), and the strong suppression – by many orders of magnitude – of the accretion rate of gas on to satellites shown by Van de Voort et al. (2017) in EAGLE. The much reduced gas fraction of such satellites then also implies that their interstellar medium rapidly increases in metallicity (Bahé et al. 2017) – a testable prediction of the scenario.

Another way to demonstrate the bias of quenched galaxies to inhabit more massive haloes is shown in Fig. 11, which plots the effective bias of galaxies as a function of stellar mass, for three different cuts in specific SFR: low (purple), intermediate (red) and high (blue). The sample cuts are the same as in Figs 8 and 10 and are listed in Table 3. [A colour version of this figure is available in the online version.]

5 COMPARISON WITH OTHER MODELS

In this section, we compare the EAGLE clustering results to two sets of models: (i) two incarnations of the GALFORM semi-analytical model, namely the version of Gonzalez-Perez et al. (2014) (hereafter GP14), and an early version of Lacey et al. (2016) (hereafter L14), both of which were used in the GAMA clustering study of Farrow et al. (2015); and (ii) the ILLUSTRIS hydrodynamical simulation described by Vogelsberger et al. (2014).

The GALFORM model assumes that galaxies form in dark matter haloes, and it uses analytical prescriptions to describe galaxy formation processes. These phenomenological prescriptions have free parameters controlling different physical processes necessary for the model to be realistic and which are tuned to fit a set of observational constraints at low redshift. GP14 and L14 use halo merger trees from the MILLENNIUM-MR7 simulation (Guo et al. 2013), which uses cosmological parameters set by WMAP7 (Komatsu et al. 2011). The G14 model has the same physical prescriptions as the Lagos et al. (2012) model, but set in a different cosmology from that used by Lagos et al. (2012), resulting in the need to re-tune the free parameters as described by G14. The main differences between GP14 and L14 are as follows: (i) the assumed stellar IMF, with GP14 using a Kennicutt (1983) IMF, while L14 switches from this IMF to a top-heavy IMF in star-bursting galaxies (see L14, for further details); (ii) the treatment of merging satellite galaxies, with GP14 using the Chandrasekhar dynamical friction time-scale in an isothermal sphere as given in Lacey & Cole (1993), while L14 use the Jiang et al. (2008) and Jiang, Jing & Han (2014) formula for the time-scale, which is empirically calibrated on N-body simulations to account for the tidal stripping of the accreting haloes; (iii) the assumed stellar population synthesis (SPS) model, with GP14 using an updated version of the Bruzual & Charlot (1993) SPS model, while L14 adopt the Maraston (2005) SPS model. Both GP14 and L14 were tuned to reproduce the $b_r$-band and $K$-band luminosity functions of Norberg et al. (2002b) and Cole et al. (2001), respectively. As GAMA is $r$-band-selected and as Farrow et al. (2015) analysis covered a larger redshift range not probed by those galaxy luminosity functions used to calibrate the GP14 and the L14 GALFORM models, Farrow et al. (2015) adjusted the GAMA GALFORM lightcone mocks constructed following Merson et al. (2013) to closely reproduce the GAMA $r$-band selection function (which, in turn,
Galaxy clustering in EAGLE

The ILLUSTRIS simulation suite also includes DMO runs. This enables us to compare the clustering of haloes between EAGLE-DMO and the DMO ILLUSTRIS (Fig. 12). The correlation function of ILLUSTRIS haloes is higher than that of EAGLE-DMO, both in real and redshift space (except for the smallest scales plotted), with the difference consistent with sample variance as judged from the scatter obtained from EAGLE-like simulation sub-volumes extracted from the significantly larger P-MILLENIUM. As discussed before, lack of large-scale power in the smaller boxes and the absence of integral constraint corrections on the clustering estimate cause the correlation functions to drop below that of P-MILLENIUM on larger scales.

Given that the EAGLE and ILLUSTRIS dark matter halo functions are very similar, whereas their galaxy stellar mass function are not (see Fig. 5 of Schaye et al. 2015), we expect some differences between the simulations in how galaxies populate the underlying dark matter haloes. This is indeed borne out by Fig. 13: ILLUSTRIS galaxies of a given stellar mass prefer lower mass haloes, by about 0.2 dex.

In Fig. 14, we compare the two-point correlation function, \( w_p \), from EAGLE galaxies with the results from ILLUSTRIS, the GP14 and L14 GALFORM models, and the GAMA survey, in four stellar mass bins. We divided \( w_p \) by the reference model \( w_p^{ref} \) of equation (7) to decrease the dynamic range in the plot. The shape of the correlation functions of EAGLE and ILLUSTRIS are very similar (we do not plot jackknife errors on the ILLUSTRIS curves to avoid clutter, but they are nearly identical to those of EAGLE), but with ILLUSTRIS offset to smaller values except for the lowest bin in stellar mass (the top panel). The poor sampling of large-scale modes in both hydrodynamical simulations, combined with the integral constraint, may lead to net offset of \( w_p \) – as we demonstrated explicitly in Fig. 2 for the dark matter haloes. The level of the offset is consistent with sample variance – as shown by the comparison to the P-MILLENIUM results. However, somewhat surprisingly, whereas haloes in ILLUSTRIS are more strongly clustered than those in EAGLE (red line above green line in Fig. 12), ILLUSTRIS galaxies are less strongly clustered than those in EAGLE at given \( M_\star \) (Fig. 14). This is related to the differences in stellar mass function: The galaxy number density is higher in ILLUSTRIS compared to EAGLE; therefore, ILLUSTRIS galaxies of given \( M_\star \) inhabit haloes of lower mass (Fig. 13), which are less clustered. This effect is relatively small, however, and we conclude that the clustering is consistent in both models, given the relatively large jackknife errors.

We can partially compensate for differences in the stellar mass function of EAGLE and ILLUSTRIS by comparing galaxy clustering at a given number density, rather than stellar mass. To do so, we select all EAGLE galaxies with \( n_{gal} < 10^{0.726} h^{-2} M_\odot \), yielding a galaxy number density of \( n_{gal} \approx 1.49 \times 10^{-2} h^3 \text{Mpc}^{-3} \) and then find the corresponding stellar mass range of \( M_\star > 10^{0.768} h^{-2} M_\odot \) in ILLUSTRIS above which the number density is equal to \( n_{gal} \). The clustering of these two samples is compared in Fig. 15, with ILLUSTRIS in red and EAGLE in green, with sample variance in the clustering of dark matter haloes with the same number density estimated by sub-sampling P-MILLENIUM volumes in grey. Selected in this way, the clustering in EAGLE is higher than in ILLUSTRIS, but the differences are consistent with sample variance.

We now turn to the semi-analytical models plotted in Fig. 14. The GALFORM model shown are the average of 26 GAMA light cone mocks as described in Farrow et al. (2015). Like in that paper, we assume that the errors on those model clustering results are negligible compared to the errors measured on the GAMA sample. See Farrow

\[\Omega_m = 0.2726, \Omega_{\Lambda} = 0.7274, \Omega_b = 0.0456, \sigma_8 = 0.809, n_s = 0.963 \text{ and } H_0 = 100 h \text{km s}^{-1} \text{Mpc}^{-1} \text{ with } h = 0.704.\]

\[^4\text{http://www.illustris-project.org}\]
et al. (2015) for a quantitative description of how adequate the GP14 and L14 GALFORM models are in describing the observed GAMA clustering. The correlation functions of GP14 and L14 are very similar, except in the second panel from the top where the GP14 model is above that of L14 on scales below $r_p \sim 1 \, h^{-1} \text{Mpc}$. Both models show stronger clustering than observed below $r_p \sim 1 \, h^{-1} \text{Mpc}$. As discussed by Farrow et al. (2015), the high values of $w_p$ in the GALFORM models are caused by an excess of satellites galaxies and/or their radial distribution in clusters. The satellite merging scheme is the principal mechanism that impacts directly the number of satellites within haloes. The two versions of GALFORM we use include the default scheme in which satellite galaxies merge on to the central galaxy after an analytically determined dynamical friction merger time-scale. Campbell et al. (2015) compared the standard GALFORM scheme with a different one (Simha & Cole 2016), in which the merger time-scale makes use of the information from dark matter sub-halo of each satellite galaxy. They find that the new scheme reduces the amplitude at small scales and shows good agreement with observational results. McCullagh et al. (2017) and Gonzalez-Perez et al. (in preparation) show that implementing the Simha & Cole (2016) merger scheme within the GP14 model and applying it to the P-MILLENIUM simulation results in clustering measurements that are in significantly better agreement with the observed ones on small scales. More detailed studies of the GALFORM clustering predictions on small scales are needed to address the known limitations of samples split by colour (see e.g. fig. 14 of Farrow et al. 2015). At large scales, the semi-analytic models show a good agreement with observational data.

To summarize, we find that both EAGLE and ILLUSTRIS reproduce the clustering of galaxies in GAMA on scales below $r_p \sim 4 \, h^{-1} \text{Mpc},

**Figure 13.** HOD of galaxies at $z = 0.1$ in four different stellar mass bins, as indicated in each panel, comparing EAGLE (green) and ILLUSTRIS (red). With the exception of the lowest stellar mass bin, galaxies of a given mass tend to reside in lower mass haloes in ILLUSTRIS compared to EAGLE. [A colour version of this figure is available in the online version.]

**Figure 14.** Comparison of galaxy clustering between different models and GAMA, as a function of stellar mass ($M_\star$ increases from the top to bottom). Different panels show the projected correlation function, $w_p(r_p)$, divided by the reference fit of equation (7). The models shown are EAGLE (green curve with jackknife errors) ILLUSTRIS (red curve – jackknife errors are not shown but are similar to those of EAGLE), GP14 (dark blue curve) and L14 (cyan curve). The observed correlation function from GAMA is shown in black with grey shading encompassing jackknife errors. The latter three sets of results are from Farrow et al. (2015). Models GP14 and L14 and GAMA curves are taken from Farrow et al. (2015). [A colour version of this figure is available in the online version.]
6 CONCLUSIONS

We have studied the two-point correlation function of galaxies at $z = 0.1$ in the **EAGLE** cosmological hydrodynamical simulation (Schaye et al. 2015). The sub-grid parameters of **EAGLE** are calibrated as described by Crain et al. (2015) to the present-day galaxy clustering properties of **EAGLE** compared to GAMA, with differences consistent with the data, given the relatively large jackknife errors and the effect of missing large-scale power.

(i) At a given stellar mass, **red** **EAGLE** galaxies are more strongly clustered than blue galaxies. In **EAGLE**, red galaxies are either satellites – with ram-pressure stripping of gas (Bahé & McCarthy 2015) and a reduction in the cosmological accretion rate of fresh gas (Van de Voort et al. 2017) – or central black hole (Trayford et al. 2016). The stronger clustering of red galaxies is then a consequence of the higher halo mass they inhabit, compared to blue galaxies of the same $M_\ast$. The difference in clustering amplitude between red and blue galaxies is too strong in **EAGLE** compared to GAMA. This overabundance of red galaxies in **EAGLE** is at least partly due to lack of numerical resolution (Schaye et al. 2015; Trayford et al. 2016), although poor sampling of massive groups and clusters plays a role as well.

(ii) On small scales, low SFR galaxies cluster more strongly for all the stellar mass bins studied. This is because galaxies with low SFR at a given mass tend to be satellites, and as before, the enhanced clustering reflects that of their more massive dark matter hosts.

(iii) Larger simulation volumes are needed to make the galaxy red, or their star formation is reduced by their central black hole (Trayford et al. 2016). The stronger clustering of red galaxies is then a consequence of the higher halo mass they inhabit, compared to blue galaxies of the same $M_\ast$. The difference in clustering amplitude between red and blue galaxies is too strong in **EAGLE** compared to GAMA. This overabundance of red galaxies in **EAGLE** is at least partly due to lack of numerical resolution (Schaye et al. 2015; Trayford et al. 2016), although poor sampling of massive groups and clusters plays a role as well.

We conclude that the galaxy clustering predicted by **EAGLE** is in very good agreement with GAMA on projected scales up to $r_p \sim 4 h^{-1} \text{Mpc}$, also when galaxies are split by colour. **EAGLE** and observed galaxies therefore inhabit haloes of similar mass, and the reduction in the SFR of **EAGLE** galaxies when they become a satellite mimics that of observed galaxies. However, the limited simulation volume of the simulation yields relatively large jackknife errors as well as large sample variance. Better tests of the realism of **EAGLE** require clustering studies in somewhat large volumes.

Comparing to other models, we find that both the GP14 and L14 semi-analytical models overestimate the galaxy clustering amplitude of galaxies. This is because the models reproduce the dominant effects of environment on galaxies, but do not focus on the details of galaxy formation itself. The models also fail to reproduce the observed clustering amplitude of red galaxies, which is a strong indicator of the relative mass of the dark matter haloes in which galaxies are found.

Figure 15. Clustering of the most massive galaxies corresponding to a mean number density of $n_{gal} \sim 1.49 \times 10^{-5} h^3 \text{Mpc}^{-3}$ at $z = 0.1$. The projected correlation function $w_p$ is divided by the reference model $w_p^{ref}$ of equation (7). The **EAGLE** and **ILLUSTRIS** simulations are plotted in green and red, respectively. Sample variance in the clustering of dark matter haloes, mass-ranked and selected to have the same mean number density, estimated using from sub-sampling the **P-MILLENIUM** volume, are shown by the grey shaded area (computed with equation 9, and including the effective bias of the sample).
amplitude at small scales, $r_s \lesssim 1\, h^{-1}\, \text{Mpc}$, while showing good agreement with GAMA on larger scales. We speculate that the excess at small scales is caused by the satellite-merging schemes implemented, which are crucial and impact directly the number of satellites and their radial distribution (Contreras et al. 2013; Campbell et al. 2015). The ILLUSTRIS simulation yields very similar clustering measures to EAGLE. At a given stellar mass, the clustering amplitude of ILLUSTRIS is lower than in EAGLE, although the difference is consistent, given the jackknife error estimates. This good agreement is slightly fortuitous: The fact that ILLUSTRIS galaxies tend to inhabit haloes of lower mass than EAGLE galaxies (consistent with ILLUSTRIS over predicting the galaxy stellar mass function for most values of $M_*$; Vogelsberger et al. 2014) – which would yield lower clustering – is partially compensated by ILLUSTRIS haloes clustering more strongly than EAGLE haloes (with the difference consistent with sample variance).

Galaxy clustering measurements provide powerful constraints on galaxy formation models. Here, we have shown that the EAGLE simulation reproduces the spatial distribution of galaxies measured in the GAMA survey, even when galaxies are split by stellar mass, luminosity and colour. This increases our confidence in the realism of the simulation. However, sample variance is still relatively large, given the small volume simulated, and better constraints require larger simulations, even when studying clustering on the smaller scales where the galaxy formation modelling is tested most stringently.

ACKNOWLEDGEMENTS

The authors would like to thank the referee, Manodeep Sinha, and Joop Schaye and Robert Crain for useful discussions. MCA acknowledges support from the European Commission’s Framework Programme 7, through the Marie Curie International Research Staff Exchange Scheme LACEGAL (PIRSES–GA–2010–269264). This work has been partially supported by PICT Raices 2011/959 of Ministry of Science (Argentina). Iz is supported by NSF grant AST-1612085 and by a CWRU Faculty Seed Grant. MCA, SP and Iz acknowledge the hospitality of the ICC at Durham, where this project was started. This work was supported by the Science and Technology Facilities Council (grant number ST/F001166/1), by the Interuniversity Attraction Poles Programme initiated by the Belgian Science Policy Office (AP P7/08 CHARM). We used the DiRAC Data Centric system at Durham University, operated by the Institute for Computational Cosmology on behalf of the STFC DiRAC HPC Facility (www.dirac.ac.uk). This equipment was funded by BIS National E-Infrastructure capital grant ST/K00042X/1, STFC capital grant ST/H008519/1 and STFC DiRAC is part of the National E-Infrastructure. PN acknowledges the support of the Royal Society through the award of a University Research Fellowship and the European Research Council, through receipt of a Starting Grant (DEGAS-259586). TT, PN, RGB and MS acknowledge the support of the Science and Technology Facilities Council (ST/L00075X/1 and ST/P000451/1). The data used in the work are publically available in the EAGLE data base described by McAlpine et al. (2016).

GAMA is a joint European–Australasian project based around a spectroscopic campaign using the Anglo-Australian Telescope. The GAMA input catalogue is based on data taken from the Sloan Digital Sky Survey and the UKIRT Infrared Deep Sky Survey. Complementary imaging of the GAMA regions is being obtained by a number of independent survey programmes including GALEX MIS, VST KiDS, VISTA VIKING, WISE, Herschel-ATLAS, GMRT and ASKAP providing UV to radio coverage. GAMA is funded by the STFC (UK), the ARC (Australia), the AAO and the participating institutions. The GAMA website is http://www.gama-survey.org/.

REFERENCES

Abazajian K. N. et al., 2009, ApJS, 182, 543
Bagla J. S., Prasad J., 2006, MNRAS, 370, 993
Bagla J. S., Ray S., 2005, MNRAS, 358, 1076
Bahé Y. M., McCarthy I. G., 2015, MNRAS, 447, 969
Bahé Y. M., Schaye J., Crain R. A., McCarthy I. G., Bower R. G., Theuns T., Mcgee S. L., Trayford J. W., 2015, MNRAS, 464, 508
Baldry I. K. et al., 2010, MNRAS, 404, 86
Baldry I. K. et al., 2012, MNRAS, 421, 621
Baldry I. K. et al., 2014, MNRAS, 441, 2440
Baugh C. M., 2006, Rep. Prog. Phys., 69, 3101
Berf lend A. A., Weinberg D. H., 2002, ApJ, 575, 587
Booth C. M., Schaye J., 2009, MNRAS, 398, 53
Bruzual A. G., Charlot S., 1993, ApJ, 405, 538
Bruzual G., Charlot S., 2003, MNRAS, 344, 1000
Calzetti D., Armus L., Bohlin R. C., Kinney A. L., Koornneef J., Storchi-Bergmann T., 2000, ApJ, 533, 682
Campbell D. J. R. et al., 2015, MNRAS, 452, 852
Camps P., Trayford J. W., Baes M., Theuns T., Schaller M., Schaye J., 2016, MNRAS, 462, 1057
Chabrier G., 2003, PASP, 115, 763
Charlot S., Fall S. M., 2000, ApJ, 539, 718
Coil A. L., Newman J. A., Cooper M. C., Davis M., Faber S. M., Koo D. C., Willmer C. A. N., 2006, ApJ, 644, 671
Coil A. L. et al., 2008, ApJ, 672, 153
Cole S., 2011, MNRAS, 416, 739
Cole S., Aragon-Salamanca A., Frenk C. S., Navarro J. F., Zepf S. E., 1994, MNRAS, 271, 781
Cole S. et al., 2001, MNRAS, 326, 255
Cole S. et al., 2005, MNRAS, 362, 505
Conroy C., Wechsler R. H., Kravtsov A. V., 2006, ApJ, 647, 201
Contreras S., Baugh C. M., Norberg P., Padilla N., 2013, MNRAS, 432, 2717
Cooray A., 2002, ApJ, 576, L105
Coupon J. et al., 2012, A&A, 542, A5
Crain R. A. et al., 2009, MNRAS, 399, 1773
Crain R. A. et al., 2015, MNRAS, 450, 1937
Crain R. A. et al., 2017, MNRAS, 474, 4204
Croton D. J., Norberg P., Gaztañaga E., Baugh C. M., 2007, MNRAS, 379, 1562
Dalla Vecchia C., Schaye J., 2012, MNRAS, 426, 140
Davis M., Peebles P. J. E., 1983, ApJ, 267, 465
Davis M., Efstathiou G., Frenk C. S., White S. D. M., 1985, ApJ, 292, 371
Dolag K., Borgani S., Murante G., Springel V., 2009, MNRAS, 399, 497
Driver S. P. et al., 2011, MNRAS, 413, 971
Eisenstein D. J. et al., 2005, ApJ, 633, 560
Farrow D. J. et al., 2015, MNRAS, 454, 2120
Favole G., McBride C. K., Eisenstein D. J., Prada F., Swanson M. E., Chuang C.-H., Schneider D. P., 2016, MNRAS, 462, 2218
Furlong M. et al., 2015, MNRAS, 450, 4486
Gonzalez-Perez V., Lacey C. G., Baugh C. M., Lagos C. D. P., Helly J., Campbell D. R., Mitchell P. D., 2014, MNRAS, 439, 264 (G14)
Goto T., Yamauchi C., Fujita Y., Okamura S., Sekiguchi M., Smail I., Bernardi M., Gomez P. L., 2003, MNRAS, 346, 601
Guo H. et al., 2013, ApJ, 767, 122
Guo H. et al., 2014, MNRAS, 441, 2398
Haardt F., Madau P., 2001, in Neumann D. M., Tran J. T. V., eds, Clusters of Galaxies and the High Redshift Universe Observed in X-rays. Available at: http://moriond.in2p3.fr
Hellwing W. A., Schaller M., Frenk C. S., Theuns T., Schaye J., Bower R. G., Crain R. A., 2016, MNRAS, 461, L11
Hinshaw G. et al., 2013, ApJS, 208, 19

MNRS 470, 1771–1787 (2017)
