Parameter inference with non-linear galaxy clustering: accounting for theoretical uncertainties

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ABSTRACT

We implement EuclidEmulator (version 1), an emulator for the non-linear correction of the matter power spectrum, into the Markov chain Monte Carlo (MCMC) forecasting code MontePython. We compare the performance of HALOFIT, HMCode, and EuclidEmulator\textsuperscript{1}, both at the level of power spectrum prediction and at the level of posterior probability distributions of the cosmological parameters, for different cosmological models and different galaxy power spectrum wave number cut-offs. We confirm that the choice of the power spectrum predictor has a non-negligible effect on the computed sensitivities when doing cosmological parameter forecasting, even for a conservative wave number cut-off of 0.2 h Mpc\(^{-1}\). We find that EuclidEmulator\textsuperscript{1} is on average up to 17\% more sensitive to the cosmological parameters than the other two codes, with the most significant improvements being for the Hubble parameter of up to 42\% and the equation of state of dark energy of up to 26\%, depending on the case. In addition, we point out that the choice of the power spectrum predictor contributes to the risk of computing a significantly biased mean cosmology when doing parameter estimations. For the four tested scenarios we find biases, averaged over the cosmological parameters, of between 0.5 and 2\(\sigma\) (from below 1\(\sigma\) up to 6\(\sigma\) for individual parameters). This paper provides a proof of concept that this risk can be mitigated by taking a well-tailored theoretical uncertainty into account as this allows to reduce the bias by a factor of 2 to 5, depending on the case under consideration, while keeping posterior credibility contours small; the standard deviations are amplified by a factor of \(\lesssim 1.4\) in all cases.

Key words: cosmology: cosmological parameters – cosmology: large-scale structure of Universe – methods: numerical – methods: statistical

1 INTRODUCTION

After very successful missions such as Planck and SDSS, many ground and space based experiments are being conducted or planned in order to learn more about our Universe. The Dark Energy Spectroscopic Instrument\textsuperscript{3} (DESI, DESI Collaboration 2016), the European space mission Euclid\textsuperscript{2} (Laureijs et al. 2011), the Legacy Survey of Space and Time (LSST) of the Vera C. Rubin observatory\textsuperscript{3} (LSST Science Collaboration 2009), the Nancy Grace Roman Space Telescope\textsuperscript{4} (Akeson et al. 2019), and the Square-Kilometre Array\textsuperscript{5} (SKA, Maartens et al. 2015) are a few well-known examples. They will provide highly accurate measurements of various cosmological observables, giving insight into the nature of dark matter and dark energy, and are likely to measure the total mass of neutrinos.

In order to exploit this wealth of high-quality data that is being produced, it has to be met by equally accurate theoretical predictions. We need to quantify the process of structure formation at high precision and to include all systematics related to the observations. Only if the highly demanding requirements on both sides are met, do we optimize our chances to determine the cosmological parameters with high enough accuracy to potentially reveal new physics.

One approach is extended perturbation theories that are able to accurately model the mildly non-linear scales. See e.g. CLASS-PT (Chudaykin et al. 2020), which was recently used for Euclid forecasts (Chudaykin & Ivanov 2019) and BOSS data analyses (Ivanov et al. 2020a,b), and EFTofLSS (publicly available as PyBird, D’Amico et al. 2021), also recently used for BOSS data analyses (Colas et al. 2020; D’Amico et al. 2020). These approaches allow us to perturbatively compute an accurate galaxy power spec-
trum to a wave number of about $k = 0.3 \, h \, \text{Mpc}^{-1}$, exceeding the range of reliability of linear perturbation theory.

However, future surveys are expected to probe much smaller scales, e.g. Euclid will collect accurate data up to $k = 10 \, h \, \text{Mpc}^{-1}$. In order to accurately model increasingly non-linear scales, those methods either need drastic improvements, or we need to use a different approach, e.g. other non-linear modelling techniques such as emulation. Cosmic emulators do not suffer from the neuralgic problems at high $k$ values as do perturbation theoretical approaches, nor do they require expensive N-body simulations beyond those used initially to train the emulator. Such emulators have received significant attention in recent years with the development of new publicly available tools and codes (Heitmann et al. 2009, 2010; Lawrence et al. 2010; Heitmann et al. 2014, 2016; Mead et al. 2015, 2016; Lawrence et al. 2017; DeRose et al. 2019; McClintock et al. 2019b; Zhai et al. 2019; Nishimichi et al. 2019; Rogers et al. 2019; Valcin et al. 2019; Bird et al. 2019; Winther et al. 2019; McClintock et al. 2019a; Fluri et al. 2019; Cataneo et al. 2019; Angulo et al. 2021; Mead et al. 2021).

We present an improvement on prediction accuracy of galaxy clustering observables thanks to integration of the publicly available EuclidEmulator1 (Euclid Collaboration: Knabenhans et al. 2019) into the MCMC sampling package MontePython6 (Audren et al. 2013b; Brinckmann & Lesgourgues 2019). EuclidEmulator1 is a cosmic emulator that quickly and accurately predicts the non-linear correction factor of the matter power spectrum for cold dark matter cosmologies with a varying time-independent Dark energy (DE) equation of state (EoS parameter $w_0$ (these are commonly referred to as “$w_0$CDM” cosmologies).

Popular methods for including non-linear information of cosmic structure formation in an MCMC run are via HALOFIT (Smith et al. 2003; Takahashi et al. 2012) or HMCode8 (Mead et al. 2015, 2016). In this work we compare these two codes to a cosmic emulator, EuclidEmulator1, which is able to capture non-linear features, e.g. non-linear damping of the BAO, as accurately as 0.3% or less compared to full N-body simulations over a wide range of both redshifts and $k$ modes – an accuracy level unprecedented by any other cosmic emulator.

The most widely used forecasting method is Fisher matrix forecasts (e.g. Tegmark et al. 1997). In its traditional form, it is conceptually very simple, but is restricted to certain very idealistic mathematical properties of the corresponding likelihood function (efforts in development of methods trying to overcome this limitation can, e.g., be found in Sellentin et al. 2014). Further, it involves computations of numerical derivatives that are likely to be unstable. For a discussion of these problems and efforts towards solutions we refer to Euclid Collaboration: Blanchard A., et al. (2020); Yahia-Cherif et al. (2021); Bhandari et al. (2021).

We connect measurements and theory in a Bayesian forecasting process through MCMC sampling method (Christensen et al. 2001; Lewis & Bridle 2002; Perotto et al. 2006), that estimates the posterior probability distributions of cosmological parameters. This is a more data driven, iterative approach, that has the advantage of not relying on derivatives nor on any assumptions of the shape of the underlying likelihood. This is the most common approach for cosmological parameter inference. Since it provides more reliable results in parameter sensitivity forecasts than classical Fisher analyses, we choose to use the MCMC method in the analysis presented in this work.

In this work, we investigate the influence of different code implementations (with different strength and weaknesses) on sensitivity forecast and parameter estimation results and we argue that using theoretical uncertainties can help deal with that (also beyond forecasts in which they are used already today). We choose to perform this analysis with the following three codes: HALOFIT and HMCode because they are very established in the field and EuclidEmulator1 as it is a representative of cosmic emulators which, as described above, have become more popular recently. We shall stress, however, that the key take-away message of this work is of qualitative rather than quantitative nature and that these qualitative results are ultimately independent of the set of codes used. Along the way of our investigations we address four main questions:

(i) What are the performance differences of EuclidEmulator1, HALOFIT and HMCode?

(ii) What is the added value of considering (mildly) nonlinear scales in parameter forecasts?

(iii) What is the impact of the choice of the (non-linear) predictor model on the parameter estimation result?

(iv) How are the forecasting results affected by different choices of theoretical uncertainty models?

In order to understand (i), we first compare the codes at the level of the matter power spectrum in section 2.1 (see Fig. 1 and 2) before performing a thorough forecast analysis in section 4.

Traditionally, in galaxy clustering analyses information from non-linear scales is removed as being "too difficult" to model correctly, and usually only information from linear and quasi-linear scales is included. However, a wealth of information is available on non-linear scales that may improve constraints. Indeed, a few parameters exhibit very clear signals only in the (mildly) non-linear regime, particularly the DE EoS parameter $w_0$, which is illustrated in Fig. 2. As such, we address question (ii) in section 4.2 by increasing the maximum wavenumber included in the analysis. This makes a case for cosmic emulators as the accuracy at nonlinear scales is one of their key selling points in contrast to models based on perturbation theory (PT).

However, once we go beyond linear scales, theoretical modelling uncertainties grow substantially, as different simulation codes increasingly disagree with each other. While the gravity-only calculation is under control up to high $k$-modes (Schneider et al. 2016; Garrison et al. 2019), baryonic feedback effects are known to affect the matter power spectrum beyond $k \sim 0.1 - 1 \, h/Mpc$ but remain poorly constrained (see e.g. Van Daalen et al. 2011; Schneider et al. 2016; Chisari et al. 2019; Aricò et al. 2020; Mead et al. 2021; Giri & Schneider 2021).

Sprenger et al. (2019) introduced a theoretical uncertainty increasing with wavenumber in order to account for the growing uncertainties at small cosmological scales, thereby making a step in

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6 We use version 1 of EuclidEmulator, referred to as EuclidEmulator1, which is available at https://github.com/miknah/EuclidEmulator. Note that during the preparation of this manuscript an improved version, EuclidEmulator2 (Euclid Collaboration: Knabenhans et al. 2021), has been made publicly available at https://github.com/miknah/EuclidEmulator2. The latter features additional cosmological parameters, namely the neutrino mass sum ($\sum m_\nu$) and dynamical dark energy ($w_a, w_0$), as well as an expanded redshift and wave number range, and an improved python wrapper.

7 See https://github.com/brinckmann/montepython_public for the latest version 3.5.

8 Since HMCode-2020 (Mead et al. 2021) is not yet available within the Boltzmann solver CLASS, we instead make use of HMCode-2016. Note that the older version is known to have difficulties in capturing non-linear features correctly, e.g. non-linear damping of the Baryon acoustic oscillations (BAO), which were improved in the newer version. While this may impact some parts of the analysis, it also illustrates the need for accurately accounting for theoretical uncertainties and problems associated with not doing so.
the direction rather than just ignoring them as was often done in the past (see Audren et al. 2013a; Baldauf et al. 2016 for earlier work in this direction). An uncertainty like that used in Sprenger et al. (2019) works well for sensitivity forecasts, where we are only interested in the uncertainties of parameter estimates and not the estimates themselves. However, results presented in this paper indicate that in the case of (mock) data analysis parameter estimation we should consider a theoretical uncertainty directly tied to the modelling choices made for the specific code that is used to fit the data, as otherwise the parameter estimates are likely to be biased.

While we discuss theoretical uncertainties and their modelling in section 2.4, they are also the focus of questions (iii) and (iv), which are addressed in section 5.1 and 5.2. Indeed, when we set up a scenario to mimic real data analysis (where the true model is not known), we find large biases in the inferred parameters, in particular when extending the analysis to larger wavenumbers. This is not unexpected, as, indeed, the three codes model non-linear scales differently. It serves as a word of caution, that when high resolution data from future surveys such as Euclid is analysed, we need to carefully account for modelling uncertainties. We make a first attempt at this using an inflated theoretical uncertainty inspired by the agreement between EuclidEmulator1 and the other two codes, and we show that with such an uncertainty envelope, we get greatly improved agreement between the codes, reducing the bias in the parameter inference without significantly decreasing the sensitivity. We leave it to future work to produce a more robust scheme for accounting for the theoretical modelling uncertainty of current non-linear parameter estimation tools.

This paper is structured as follows: In section 2 present a comparison of the three codes analyzed in this paper – EuclidEmulator1, HALOFIT, and HMCode – at the level of the matter power spectrum and discuss both known and unknown systematics impacting parameter estimations and sensitivity forecasts as well as how they are related to code differences. In this section we motivate the need for a theoretical uncertainty in order to mitigate unwanted effects due to those systematics. Next, in section 3 we introduce the analysis methodology employed in this work, including some theoretical considerations as well as the strategy of our computational experiments. The results are presented and discussed in section 4 and 5. We finally conclude this paper in section 6.

2 SYSTEMATIC UNKNOWNS IN NON-LINEAR COSMOLOGICAL STRUCTURE FORMATION

2.1 Code comparison at the matter power spectrum level

Before jumping to the MCMC results, let us first quantify differences between the three non-linear prescriptions for predicting the matter power spectrum $P(k, z; p)$ considered in this paper: HALOFIT (“HF”), HMCode (“HM”) and EuclidEmulator1 (“EE”). This will help us understand the differences we may see in the following section.

Due to differences in how the non-linear prescriptions model non-linear structure formation, we expect a difference in how each code responds to a changes in cosmological parameters. In the following, we want to compare the non-linear prescriptions to each other in order to quantify these differences. Taking EuclidEmulator1 as the reference, we vary one cosmological parameter at a time, keeping all others fixed to a fiducial cosmological model (Euclid Reference Cosmology as defined in Euclid Collaboration: Knabenhans et al. 2019). We define the model ratio

$$R_{\text{model}}(k, z; p) = \frac{P_{\text{model}}(k, z; p)}{P_{\text{EE}}(k, z; p)}$$

where model $\in \{\text{HF, HM}\}$ and $p \in \{\omega_{0}, \omega_{m}, n_{s}, h, w_{0}, A_{s}\}$.

We consider the difference with respect to the reference cosmological model

$$\Delta R_{\text{model}}(k, z; p) = R_{\text{model}}(k, z; p) - R_{\text{model}}(k, z; p_{\text{ref}})$$

This quantifies how, for a given non-linear prescription, the matter power spectrum changes when we vary one of the cosmological parameters.

In Fig. 1, we show the ratio from Eq. 1, i.e. how well HALOFIT (left panel) and HMCode (right panel) agree with EuclidEmulator1. This is shown at four redshifts ($z = [0, 0.7, 1, 2]$), which includes the ratio today and spans the redshift range of the Euclid galaxy clustering likelihood ($z = 0.7$ to $z = 2$), and for a fixed cosmology (this is also shown for a varying cosmology at $z = 1$ in Fig. 2 columns 1 and 3). We show only the redshift dependence for the reference cosmological model, as the difference from Eq. 2 changes negligibly across redshift. We see several notable trends:

- Both HALOFIT and HMCode see an overall small shift of the matter power spectrum to smaller scales (larger $k$), resulting in a disagreement on BAO scales ($k \approx [0.05, 0.5]$ $h$ Mpc$^{-1}$) as seen from the sharp wiggles and the small overall amplitude shift on those scales. For both codes, the disagreement on BAO scales compared to EuclidEmulator1 increases with smaller redshift. The disagreement on BAO scales compared to EuclidEmulator1 is obviously highly scale-dependent due to the BAO feature, but it is maximally about 5% for HALOFIT at $z = 0$ (about 2% in the Euclid likelihood redshift range) and about 8% for HMCode at $z = 0$ (about 4% in the Euclid likelihood redshift range). For HALOFIT, we see a larger amplitude of the matter power spectrum compared to EuclidEmulator1 across all redshifts considered, while for HMCode the sign of the amplitude shift depends on redshift: at $z = 2$ the amplitude is shifted to smaller values, while at the other redshifts considered the amplitude of the matter power spectrum is shifted to larger values.

- On small scales ($k \gtrsim 0.5$ $h$ Mpc$^{-1}$), we see a very different behavior between each of the three codes. Across all redshifts, HALOFIT sees a larger amplitude of the matter power spectrum compared to

![Figure 1. Redshift evolution for power spectrum ratio (as defined in Eq. 1) between HALOFIT and EuclidEmulator1 (left panel) and HMCode and EuclidEmulator1 (right panel). The evolution, which is only weakly dependent on the cosmology, is shown for the Euclid reference cosmology.](image)

9. EuclidEmulator1 does not accept $A_{s}$ as input but only $\sigma_{8}$. The conversion for each set of cosmological parameters is done with CLASS.
Figure 2. Variation of the non-linear matter power spectrum due to variation of cosmological parameters. Each row shows the variation of one cosmological parameter: $\omega_b$ in the first row, $\omega_m$ in the second, etc. All parameters that are not varied in any given row, are set to their corresponding values in the Euclid Reference Cosmology. Notice that the Euclid Reference cosmology corresponds to the middle parameter values in each of the parameter value legends, i.e. $\omega_b = 0.02172$, $\omega_m = 0.1432$, $n_s = 0.96000$, $h = 0.67000$. $A_s = 2.10 \times 10^{-9}$ is always fixed. The variation is shown for the quantity $R_{\text{model}}$ as defined in Eq. 1 (columns 1 and 3). In order to emphasize (i) which scales are affected the most by the variation of a specific parameter and (ii) how strongly $R$ is impacted by the variation of each parameter, we also plot the $\Delta R$ with respect to the $R$ of the Euclid Reference Cosmology (columns 2 and 4). It is evident, for instance, that the variation of $h$ and $w_0$ has a much larger effect on $R$ than e.g. $\omega_m$. Notice that the tested parameter values are taken to be the Euclid Reference Cosmology value (the central value) $\pm \sigma_{\text{Planck2018}}$, $i \in 1, 2$ where $\sigma_{\text{Planck2018}}$ denotes the standard deviation of each parameter $p$ of the Planck 2018 results (for more details see section 17.14 on p. 301 in Planck Collaboration 2018).

EuclidEmulator1, but in a scale-dependent and redshift-dependent way: at $z = 0$ the disagreement compared to EuclidEmulator1 peaks at 4% at $k \approx 2.5 \, h \, \text{Mpc}^{-1}$ and is about 2–3% across the rest of the interval ($k \in [0.5, 5] \, h \, \text{Mpc}^{-1}$); at $z = 0.7$ the disagreement peaks at about 3.5% at $k \approx 0.8 \, h \, \text{Mpc}^{-1}$, dropping to no difference at $k \approx 2 \, h \, \text{Mpc}^{-1}$, before rising to 3% at $k \approx 5 \, h \, \text{Mpc}^{-1}$; at $z = 1$ the disagreement peaks at about 3.5% at $k \approx 1 \, h \, \text{Mpc}^{-1}$, dropping to no difference at $k \approx 3 \, h \, \text{Mpc}^{-1}$, before rising to 1.5% at $k \approx 5 \, h \, \text{Mpc}^{-1}$; at $z = 2$ the disagreement peaks at about 7.5% at $k \approx 2 \, h \, \text{Mpc}^{-1}$, dropping to about 2–3% near either end point of the interval $k \approx [0.5, 5] \, h \, \text{Mpc}^{-1}$.

- On small scales ($k \geq 0.5 \, h \, \text{Mpc}^{-1}$) for HMCode, we see a decrease in the amplitude of the matter power spectrum compared to EuclidEmulator1 for all redshifts of about 1–3%, except for $z = 0$ and $z = 2$, where the disagreement changes from a 3% decrease in amplitude at around $k \approx [0.5, 1] \, h \, \text{Mpc}^{-1}$ to a 2–4% increase at around $k \approx [2, 5] \, h \, \text{Mpc}^{-1}$: at $z = 2$ this comes as a peak at 4% at $k \approx 2.5 \, h \, \text{Mpc}^{-1}$ that again decreases to around 1% at...
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$k = 5 \, h \, Mpc^{-1}$, while at $z = 0$ this comes as a steady increase until a peak at 4% at $k \approx 5 \, h \, Mpc^{-1}$.

In Fig. 2, we vary the cosmological parameters one at a time at redshift $z = 1$: the baryon abundance $\omega_b$ (first row), the matter abundance $\omega_m$ (second row), the tilt of the primordial power spectrum $n_s$ (third row), the Hubble parameter $h$ (fourth row), the equation of state of dark energy $w_0$ (fifth row), and the amplitude of the primordial power spectrum $A_s$. We show the ratio from Eq. 1 in columns 1 (HALOFIT) and 3 (HMCode), and the difference from Eq. 2 in columns 2 (HALOFIT) and 4 (HMCode). Notice that the tested parameter values are taken to be the Euclid Reference Cosmology value (the central value) $\pm 2\sigma_{p, Planck2018}$ for $i \in 1, 2, 3$ where $\sigma_p, Planck2018$ denotes the standard deviation of each parameter result from the Planck 2018 results (for more details see section 17.14 on p. 301 in Planck Collaboration 2018).10

We see that the difference (Eq. 2) of HALOFIT and HMCode compared to EuclidEmulator1 is fairly small for $\omega_b$, $\omega_m$, and $n_s$. For $\omega_b$, the difference peaks at about 0.5% on very large, unobservable scales (small $k$) for a $2\sigma$ variation in the parameter, with negligible difference between HALOFIT and HMCode; for $\omega_m$, the difference peaks on observable scales, but with only about 0.4% maximal difference for HMCode (the difference is smaller for HALOFIT) for a $2\sigma$ variation in the parameter; for $n_s$, the difference is largest for HALOFIT (it is much smaller for HMCode), peaking at around 0.4 – 0.6% for a $2\sigma$ variation in the parameter.

For the two parameters, $h$ and $w_0$, we see over percent level differences: for $h$ we see a large difference of HALOFIT and HMCode compared to EuclidEmulator1 across BAO and small scales, with the largest difference on BAO scales, where the shift in BAO peak location results in a difference of up to 1.5% for HMCode and 1.1% for HALOFIT; for $w_0$ the largest difference is on small scales, ranging from up to a few per mil to BAO scales to $1 - 1.5\%$ (with largest difference compared to EuclidEmulator1 being for HALOFIT) at around $k \approx 1.5 - 2 \, h \, Mpc^{-1}$.

For the last parameter, $A_s$, we see again somewhat smaller differences, that reach their maximal amplitude of about 0.6% at $k \approx 1.5 \, h \, Mpc^{-1}$ for HALOFIT. For HMCode the maximal amplitude is marginally smaller (about 0.5%) and it is reached only at $k \approx 2 \, h \, Mpc^{-1}$.

We will see in section 5 that the differences in Fig. 1 lead to parameter estimation biases when performing a mock data analysis. This will be discussed further there.

2.2 Sensitivity forecasting and parameter estimation at non-linear scales

In the past, galaxy clustering data analyses used to include a conservative non-linear cut-off in $k$-space, above which all data is discarded. As, however, it is exactly these non-linear scales that modern cosmological surveys aim to exploit, such a cut-off defeats the point of the surveys. Yet, the theory required to analyse the data from non-linear scales is based on imperfect understanding of the non-linear structure evolution in the Universe due to hydrodynamics, galaxy formation, feedback processes etc. As a consequence, the codes implementing the theory (such as HMCode, HALOFIT and EuclidEmulator1 discussed in the previous subsection) tend to increasingly disagree the more non-linear scales are considered.

We will now discuss how such theory and code disagreements can influence cosmological parameter inference on a grander scale. There are two primary tasks/goals in cosmological parameter inference:

- Measuring cosmological parameters with high accuracy (low bias)
- Measuring cosmological parameters with high precision (low variance/uncertainties) in a robust manner

For both tasks, we typically compare a given theory to (mock) data by computing (or sampling) a posterior distribution for the cosmological parameters in a Bayesian way: the parameter value is estimated by the maximum posterior probability while the uncertainties are derived from the credible contours of the posterior. As credible contours can be reduced by simply ignoring sources of uncertainty, robustness plays also a key role in precision measurements: we try to find the credible contours that are simultaneously as small and as correct as possible.

“Parameter estimation” is the task in which the values of the parameters themselves are to be determined, i.e. in which the maximum posterior probability is to be evaluated. In a real-case scenario, the data comes from an actual observation/experiment. In past decades, data was mainly collected from the realm of linear structure formation, which stands on firm theoretical grounds such that there was little disagreement about the theories used to analyze the data. However, modern cosmological experiments have entered the realm of non-linear structure formation and the corresponding data contains a wealth of valuable information which should not be disregarded. Nevertheless, non-linear structure formation is not known at the same level of precision as its linear counterpart and, hence, it is important to remember that the choice of the predictor may lead to slightly different estimates of the parameter values. In other words, the obtained parameter values are not universal but “remember” the theory/code they were estimated with. A similar result has been found and published by Martinelli et al. (2021), who used weak lensing observables for their analysis. In this paper, we offer a proof-of-concept solution for this problem inspired by work previously published: we suggest that code-dependence is best handled by introducing a “theoretical uncertainty” into the data analysis (in Audren et al. 2013a; Baldauf et al. 2016; Sprenger et al. 2019 this concept is introduced solely in the context of sensitivity forecasting). Theoretical uncertainties are discussed more deeply in section 5.2.

“Parameter sensitivity forecasting” (or short sensitivity forecasting) is a common method to project how precisely a given experiment is able to determine the cosmological parameters. As such, it provides guidance in decision making processes regarding the experimental design of a given survey and verifies to the community that the science goals are achievable.

In contrast to the task of parameter estimation, sensitivity forecasts are usually performed with mock data. This mock data is typically generated using the same code and theory as is used to compute/sample the posterior distribution, thereby we assume that the data is perfectly modelled with appropriate uncertainties. Developments in cosmological structure formation theory, and the ever increasing accuracy of computational methods able to predict cosmological observables at non-linear scales, allow scientists to better take into account systematics in their analyses. For this reason, it is an interesting exercise in its own right to measure the performance of newly developed methods applied to sensitivity forecasts.

10 Notice that setting $w_0 = w_0^{ref} - 2\sigma_p, Planck2018 = -1.2$ while keeping all other cosmological parameter fixed to their fiducial values results in a $\sigma_p$ value outside the parameter space accepted by EuclidEmulator1. For this reason the corresponding $w_0$ value was manually adjusted to $w_0 = -1.19$. 

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and benchmark them against more established tools. However, even when various sources of uncertainties are taken into account, it is highly unlikely that none have been missed at all.

In the context of sensitivity forecasts, theoretical uncertainties can hence be used in order to (1) fold in uncertainties due to neglected systematics (such as e.g. baryonic physics or scale-dependent galaxy bias) and (2) to deal with unknown uncertainties that have not been modelled. As, however, different codes implement different theories and assumptions and hence take into account different sources of uncertainties, the result of sensitivity forecast does also depend on the choice of code used. A theoretical uncertainty hence makes a sensitivity forecast more robust and less dependent on the modelling choices implemented into the codes used in the forecast.

2.3 Treatment of known systematics: parameter inference with cosmological emulators

There is a growing number of power spectrum predictors available on the market, ranging from full-fledged cosmological simulations over halo model derivatives and PT-based models all the way to purely data-driven approaches. The first kind offers the most accurate predictions but are not useful in neither have limited use for sensitivity forecasts and parameter estimations due to their immense computational expense. PT- and halo model-based models are orders of magnitudes cheaper to evaluate than simulations but they suffer from more systematics as they are based on calibrated fitting functions. In the past, they were the choice of power spectrum predictors in parameter inference thanks to their low computational cost. However, emulators are comparable with (semi-)analytical models in terms of computational expense but often carry less systematics as they are based on simulations. It must be emphasized, though, that the remaining systematics need to be treated carefully. Observations necessarily contain information whose modelling needs to be taken into account sources of uncertainty such as redshift space distortions (RSDs), the Alcock-Paczynski effect, galaxy bias, baryonic effects on structure formation etc. (see section 3.1).

Although this is not a special trait of emulators, many are actually designed to predict correction factors to deal with individual systematics. Examples are e.g. the non-linear correction of the matter power spectrum (Euclid Collaboration: Knabenhans et al. 2019, 2021), the baryonic emulator (Schneider et al. 2020; Giri & Schneider 2021), or the halo bias (Valcin et al. 2019; McClintock et al. 2019a). This way, multiple individual emulators, each modelling a different effect, can, under certain assumptions, be stacked on top of each other to take a multitude of systematics into account.

It can therefore be anticipated that future parameter inference studies will use emulators as robust predictors that may include uncertainties in a modular way. A possible approach to deal with unknown sources of systematic errors is discussed in the subsequent subsection.

2.4 Treatment of unknown systematics: parameter estimation with theoretical uncertainty

Theoretical uncertainties and their applications in the context of sensitivity forecasts have been studied in previous work (e.g. Audren et al. 2013a; Baldauf et al. 2016; Sprenger et al. 2019). In this work, we propose to extend the use of theoretical uncertainties in order to mitigate the negative impact of theory and code disagreements on parameter estimations due to unknown systematic effects (see section 2.2). To this end it is natural to start with a theoretical uncertainty readily available such as the one published in Sprenger et al. (2019) and study its effects on parameter estimation. We hence review this particular uncertainty envelope in section 2.4.1. In addition we introduce a new theoretical uncertainty in section 2.4.2 that is tailored to the power spectrum prediction codes used in this paper. The comparison of the two theoretical uncertainties is presented later in the context of a mock parameter estimation in section 5.2.

2.4.1 Review of the simulation-based theoretical uncertainty

We start out reviewing the theoretical uncertainty of Sprenger et al. (2019), which aims to identify the theoretical precision with which we will be able to analyze Euclid data using cosmological simulations. This uncertainty envelope, in the following denoted \( \sigma_{\text{th},\text{S}} \), is roughly given by the agreement between hydrodynamical simulations at the time of publishing of that work, which is expected to improve by the time Euclid data is available and being analyzed. The envelope \( \sigma_{\text{th},\text{S}} \) may be sufficient for any traditional forecasting with an aim to predict the sensitivity of future surveys as state-of-the-art tools on average disagree at a level that is mostly captured by \( \sigma_{\text{th},\text{S}} \) (Sprenger et al. 2019).

The approach of Sprenger et al. (2019) takes three uncertainty values at three different wave numbers as inputs,

- 0.33% at \( k = 0.01 \text{ h Mpc}^{-1} \),
- 1.0% at \( k = 1 \text{ h Mpc}^{-1} \),
- 10.0% at \( k = 10 \text{ h Mpc}^{-1} \)

and interpolates between these points using the ansatz for the uncertainty envelope

\[
\sigma(k, z) = \begin{cases} 
    a_1 \exp c_1 \log 10 \left( \frac{k}{h \text{ Mpc}^{-1}} \right), & \frac{k}{h \text{ Mpc}^{-1}} < 0.3 \\
    a_2 \exp c_2 \log 10 \left( \frac{k}{h \text{ Mpc}^{-1}} \right), & \frac{k}{h \text{ Mpc}^{-1}} > 0.3 
\end{cases}
\]  

(3)

where \( \sigma \approx \sigma_{\text{th},\text{S}} \) and \( k \) depends on the line of sight direction \( \mu \) and the cosmology, \( k = k(\mu, \text{cosmo}) \) (for more details on this point see Sprenger et al. 2019).

In general, we define

\[
k_{\text{th}}(z) = k_0 \cdot (1 + z)^{\frac{1}{2}},
\]  

(4)

such that \( k_{1 \text{ h Mpc}^{-1}}(z) \) is defined accordingly with \( k_0 = 1 \text{ h Mpc}^{-1} \). In order to interpolate between the points mentioned above the coefficients are set to

\[
a_1 = 0.014806, a_2 = 0.022047, c_1 = 0.75056, c_2 = 1.5120.
\]  

(5)

See Fig. 3 for an illustration of the uncertainty envelopes, including their redshift dependence. The blue lines represent \( \sigma_{\text{th},\text{S}} \) at three different redshifts.

In Sprenger et al. (2019) (see their section 5 for details) the theoretical 1\( \sigma \) uncertainty was given as

\[
\sigma_{\text{th}}(k, \mu, z) = \left. \frac{V_r(z)}{2(2\pi)^2} \frac{k^2 \Delta k \Delta \mu}{\Delta z^2} \alpha(k, \mu, z) P_{gg}(k, \mu, z) \right]^{1/2}
\]  

(6)

with redshift correlation length \( \Delta z \), survey redshift bin width \( \Delta z \), survey volume per redshift bin \( V_r(z) \), and wave number correlation length \( \Delta k \). This uncertainty is then simply added in quadrature to the observational uncertainty in the \( \chi^2 \) computation.

However, we note that the publicly available likelihood (as of MontePython version 3.4) has a bug resulting in an inverted ratio of
(\(\frac{\Delta z}{\Delta z}\))\(^{-1}\), which with the chosen redshift correlation length \(\Delta z = 1\) and survey redshift bin width \(\Delta z = 0.1\) gives rise to a factor \((100)^{1/2} = 10\) smaller uncertainty than intended. Since correlation lengths are inherently difficult to determine, and Sprenger et al. (2019) used very conservative correlation lengths, this can largely be interpreted as an adjustment to less conservative correlation lengths that would still be within the range of possible values: i.e. if we use \(\Delta z = 0.1\) instead of \(\Delta z = 1\), \(\Delta \mu = 1\) instead of \(\Delta \mu = 2\), and \(\Delta k = 0.02\ h^{-1}\) instead of \(\Delta k = 0.05\ h^{-1}\) then without the bug we would have a factor 2 smaller uncertainty than intended. Although unfortunate, given that the precision of future cosmological simulations that the uncertainty is intended to mimic is not known and Sprenger et al. (2019) instead used the precision near the time of writing, a factor 2 smaller uncertainty (improved precision) is not unrealistic.

In contrast to the hydrodynamical simulations mentioned above, HALOFIT and HMCode exhibit an error on BAO scales that is sufficient to bias the result of data analyses\(^\text{11}\). While we would expect the Euclid Collaboration analysis of future data to employ methods that accurately capture the BAO scales, leaving the Sprenger et al. (2019) theoretical uncertainty robust for forecasts, we find in this paper (cf. section 5.1) that the bias remains for analyses that employ commonly used current fast methods such as e.g. HALOFIT and HMCode. Reducing this bias motivates the construction of a new uncertainty envelope that will be the topic of the next subsection.

### 2.4.2 Theoretical uncertainty based on non-linear predictors

**EuclidEmulator1** has been compared to HALOFIT and HMCode in Fig. 2 (for a comparison between EuclidEmulator1 and HALOFIT also see Euclid Collaboration: Knaebenhans et al. 2019). The agreement at linear scales is almost perfect. At intermediate scales the agreement is at the 1-2\% level for redshift \(z = 2\), while it is at the few percent level for \(z = 0\). For smaller scales the agreement is at the few percent level for all redshifts \(0 \leq z \leq 2\). We therefore construct a new theoretical uncertainty envelope \(\sigma_{th,KB}\) based on the disagreement between these codes (see green lines in Fig. 3).

We find that this new uncertainty envelope reduces the bias on (mock) data analyses significantly (discussion in section 5.2). However, while the disagreement between HALOFIT/HMCode and EuclidEmulator1 oscillates considerably in \(k\) at intermediate and small scales for all \(z\), for the sake of simplicity we chose to model the uncertainties with a constant close to the maximal disagreement found on scales \(0.1\ h\ Mpc^{-1} \leq k \leq 1\ h\ Mpc^{-1}\). Hence, at these scales our uncertainty is somewhat overly conservative. The resulting uncertainty envelope follows the one by Sprenger et al. (2019) up to \(k = 0.05\ h\ Mpc^{-1}\) and is fixed to 5\% above \(k = 0.15\ h\ Mpc^{-1}\) at \(z = 0\). The intermediate part is modelled by Eq. 3 using the coefficients \(a_2 = 0.2998\) and \(c_2 = 2.2517\), with the amplitude fixed to \(\sigma_{th,KB}\) at \(k = 0.05\ h\ Mpc^{-1}\).

Another question is up to which \(k_{max}\) mode these uncertainties can be trusted. Such a \(k_{max}\) represents a cut-off scale above which the theoretical uncertainty is de facto infinite. In other words, all information coming from \(k\) modes larger than \(k_{max}\) is simply discarded. We use two different cut-off points: one ultra-conservative disregarding all scales above \(k_{max} = k_{th,KB} = 0.2\ h\ Mpc^{-1}\) at \(z = 0\), and one informed by our choice of galaxy bias model (see section 3.1), which is only reliable up to about \(k_{max} = k_{th,KB} \approx 0.4\ h\ Mpc^{-1}\) at \(z = 0\). Both scale with redshift as defined in Eq. 4.

\[^{11}\text{Note this specific issue was mitigated in the latest version of HMCode: Mead et al. (2021).}\]

---

Figure 3. The two different envelope functions \(\alpha\) for theoretical uncertainty as defined in Eq. 3. The solid, dashed and dotted lines indicate the uncertainty envelope at redshifts \(z = 0\), \(z = 0.7\) and \(z = 2\), respectively, where \([0, 0.7, 2]\) is the redshift range of the Euclid likelihood. Blue lines indicate the original theoretical uncertainty \(\sigma_{th,KB}\) from Sprenger et al. (2019) and green lines correspond to \(\sigma_{th,KB}\), a rough estimate of the agreement between EuclidEmulator1 and HMCode, as described in the text in section 2.4. The vertical dashed green line indicates \(k = 0.2\ h\ Mpc^{-1}\) which is a popular choice for the upper \(k\) cut-off value. In most cases, we use the theoretical uncertainty \(\sigma_{th,KB}\), but at the end we compare to the other uncertainty for a cross-comparison case.

### 3 ANALYSIS METHODOLOGY

In section 1, we listed the four key questions that guide our efforts in this paper:

(i) What are the performance differences of **EuclidEmulator1**, HALOFIT and HMCode?

(ii) What is the added value of considering (mildly) nonlinear scales in parameter forecasts?

(iii) What is the impact of the choice of the (non-linear) predictor model on the parameter estimation result?

(iv) How are the forecasting results affected by different choices of theoretical uncertainty models?

In order to answer these questions we run and analyze a set of MCMCs. In this section we discuss the general aspects of the MCMC analysis methodology as well as the experimental setup used to generate those MCMCs. The results and interpretations will be reported later in section 4 and 5.

### 3.1 Modelling the galaxy-galaxy power spectra

**EuclidEmulator1** predicts the non-linear correction to the dark matter power spectrum only, which is not an observable quantity. A related quantity that we can measure from observations is the galaxy-galaxy power spectrum \(P_{gg}(k, \mu, z)\), where \(\mu\) is the cosine between the wave mode \(k\) and the line of sight direction \(e_x\). Following the expansion in Sprenger et al. (2019) and using their notation we have

\[
P_{gg}(k, \mu, z) = f_{AP}(z) \times f_{res}(k, \mu, z) \times f_{RSD}(\hat{k}, \hat{\mu}, z) \times b^2(z) \times P_{\delta\delta}(k, z)
\]

where \(f_{AP}\) denotes the correction due to the Alcock-Paczynski effect, \(f_{res}\) is a correction factor necessary because the resolution is limited, \(f_{RSD}\) accounts for the RSDs and \(b(z)\) is the galaxy bias taking
into account that galaxies are only biased tracers of the underlying matter distribution. Notice that the parameters \(k\) and \(\mu\) are actually functions of the underlying cosmology. Parameters with a hat denote that they are based on the true/real but unknown cosmology, while those without a hat are based on an assumed underlying cosmology. For a more elaborate explanation of this see Sprenger et al. (2019).

For the galaxy bias we choose a linear, redshift-dependent bias \(b(z) = \beta_0(1+z)^{\beta_1}\) as in Sprenger et al. (2019), where \(\beta_0\) and \(\beta_1\) are included as nuisance parameters in the MCMC analysis. Since galaxy bias is known to be better described by a scale-dependent galaxy bias (e.g. Desjacques et al. 2018; Giusarma et al. 2018), we choose to cut the analysis at an ultra-conservative (e.g. Desjacques et al. 2018; Giusarma et al. 2018), we choose to cut the analysis at an ultra-conservative scale-dependent bias to these scales would be picked up by the theoretical uncertainty.

The peculiar motion of galaxies give rise to the so-called Fingers of God (Tully & Fisher 1978; Jackson 1972) and are treated separately as an exponential suppression of power following Bull et al. (2015), where the suppression is governed by a nuisance parameter \(c_{\eta_0}\), that we marginalize over. The remaining RSD effect is described by the Kaiser formula (Kaiser 1987), which relates the galaxy power spectrum \(P_{\delta\delta}\) to the dark matter power spectrum \(P_{\delta\delta}^{\text{linear}}\). EuclidEmulator1 directly gives us an nonlinear correction (NLC) factor \(B(k,z) = 1 + 2\mu^2f + \mu^4f^2\) \(P_{\delta\delta}^{\text{linear}}(k,z)\), which, setting aside the other correction factors for the moment, allows us to write the RSD and bias correction as (see section A)

\[
P_{\delta\delta}(k,\mu,z) = b^2(z) \left[ B(k,z) + 2\mu^2f + \mu^4f^2 \right] P_{\delta\delta}^{\text{linear}}(k,z). \tag{8}
\]

For HALOFIT and HMCCode, which return a non-linear power spectrum, we find identical results expressing this as (see section A)

\[
P_{\delta\delta}(k,\mu,z) = b^2(z) \left[ 1 + \mu^2f \right] ^2 \rho_{\delta\delta}^{\text{non-linear}}(k,z) \tag{9}
\]

The other correction factors are multiplicative and can trivially be included in these expressions. For deeper discussion of all of these correction factors we refer the reader to Sprenger et al. (2019) and references therein.

### 3.2 Experimental setup

We run the MCMCs with MontePython version 3.1 (Audren et al. 2013b; Brinckmann & Lesgourgues 2019) and CLASS version 2.8 (Blas et al. 2011; Lesgourgues 2011). We use the publicly available Euclid galaxy clustering likelihood\(^\text{12}\) from Sprenger et al. (2019), modified according to section 3.1 and section 2.4 to allow for including non-linear corrections by the EuclidEmulator1 and a modified theoretical uncertainty scheme. The likelihood uses 13 redshift bins in the redshift range \(0.7 < z < 2\), a sky fraction \(f_{\text{sky}} = 0.3636\), nine \(\mu\) bins, and 100 logarithmically spaced bins in \(k\)-space from \(k_{\text{min}} = 0.02\) Mpc\(^{-1}\) to \(k_{\text{max}} = 0.2\) or 0.4 Mpc\(^{-1}\) (see the end of section 2.4 for a discussion of \(k_{\text{max}}\)). For more details we refer the reader to Sprenger et al. (2019).

We combine the Euclid likelihood with the fake Planck likelihood and noise spectra from Brinckmann et al. (2019), which is based on the old “fake planck bluebook” likelihood with specifications updated to the full mission and omitting the 217 GHz channel, in order to avoid foreground complications. The end result is a likelihood in good agreement with the sensitivity of the Planck 2018 results (Planck Collaboration 2020), but with a slightly better sensitivity for the optical depth to reionization, \(\sigma(\tau_{\text{reio}})\). The fake Planck likelihood considers a maximum multipole of \(\ell_{\text{max}} = 3000\) for temperature, polarisation and lensing, with the noise exploding long before the maximum multipole (see Figure 1 of Brinckmann et al. 2019).

### Table 1: Cosmological (top panel) and nuisance parameters of the fiducial cosmology used for all MCMC forecasting runs. The middle panel lists the varying RSD and galaxy bias nuisance parameters of the Euclid likelihood (see section 3.1), while the bottom panel contains baryonic feedback nuisance parameters of HMCCode, which remain fixed. The fiducial cosmology coincides with the Euclid Reference Cosmology (as defined in the Euclid Collaboration: Knabenhans et al. 2019). Notice that the minimal and maximal values are broadly equal to the parameter box of EuclidEmulator1 (keep in mind that \(\Omega_{\text{cdm}}h^2\) is not a direct input parameter to EuclidEmulator1).

| Parameter | fiducial value | min | max |
|-----------|---------------|-----|-----|
| \(\Omega_b h^2\) | 0.022 | 0.0215 | 0.0235 |
| \(\Omega_{\text{cdm}} h^2\) | 0.121 | 0.1091 | 0.1311 |
| \(n_s\) | 0.96 | 0.9283 | 1.0027 |
| \(h\) | 0.67 | 0.6155 | 0.7307 |
| \(\sigma_8\) | 0.9619 | -1.25 | -0.75 |
| \(\tau_{\text{reio}}\) | 8.24 | 0.0 | none |
| \(\sigma_{\text{NL}}\) | 7.0 | 4.0 | 10.0 |
| \(\beta_0\) | 1.0 | none | none |
| \(\beta_1\) | 1.0 | none | none |
| \(\eta_0\) | 0.603 | fixed |
| \(\epsilon_{\min}\) | 3.13 | fixed |

### 3.3 MCMC analysis

We use the MCMC technique to compare the agreement between a fiducial and a predicted galaxy-galaxy power spectrum, \(P_{\delta\delta}\), where the latter takes a cosmology as an input argument. While \(P_{\delta\delta}\) is the observable galaxy clustering quantity used in this work, the three codes compared to one another all predict the non-observable matter power spectrum \(P_{\delta\delta}\). Technical details about how \(P_{\delta\delta}\) is related to the \(P_{\delta\delta}\) and how the NLC predicted by EuclidEmulator1 is worked into this relation are given in Eq. 7 and Eq. 8.

In this paper, we perform sensitivity forecasts and mock parameter estimations instead of real-world data analysis. Hence, we first have to construct a fiducial model. As will be explained in more detail below, our MCMC runs can be split into two classes, which we refer to as “auto-comparison” (ac) and “cross-comparison” (cc) tests (see section 5 for an explanation why we do this). In the case of auto-comparisons, the fiducial model is set up with the same cosmological model, and cut-off scale as the model used to perform the theoretical predictions during the MCMC. For cross-comparisons, on the other hand, we compute a fiducial spectrum with EuclidEmulator1 in all cases and use either HALOFIT or HMCCode for the theoretical predictions. Notice that the choice of the theoretical uncertainty model does not affect the computation of the fiducial result.

We let the MCMCs run on the zBox4 cluster based at the University of Heidelberg with 2000 MCMC steps and 40% burn-in time.

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\(^{12}\) Note that in cosmological inference the term “likelihood” is typically a package containing all modelling of observables plus the likelihood computation.
of Zurich\textsuperscript{13}. We run twelve chains simultaneously for each MCMC. For any information about how this works in detail, we refer the reader to Audren et al. (2013b); Brinckmann & Lesgourgues (2019).

To determine convergence we use the standard criterion by Gelman & Rubin (1992), $R - 1 < 0.01$, where $R$ is the Gelman-Rubin coefficient\textsuperscript{14}. We let the MCMCs run until this convergence criterion is satisfied in all parameters.

### 4 SENSITIVITY FORECASTS

Traditionally, different power spectrum predictors are compared at the power spectrum level as is done e.g. in Lawrence et al. (2010); Bird et al. (2012); Mead et al. (2016); Lawrence et al. (2017); Euclid Collaboration: Knabenhans et al. (2019); Angulo et al. (2021); Euclid Collaboration: Knabenhans et al. (2021); Mead et al. (2021). However, it is not often studied directly how the differences between the predictors impact parameter forecasts in the end (see, however, Martinelli et al. (2021) for a complementary analysis for cosmic shear). We close part of this gap in this section by analyzing the code dependence of forecasting results on the three codes studied in this paper, leaving investigations of this aspect with other codes to future work.

Code dependencies in forecasting results could manifest themselves in different ways: uncertainty contours could be smaller or show different degeneracy directions. Ultimately, we would like to understand the reasons behind any differences, which could be due to mistakes in the modelling, e.g. underestimating uncertainties, not accurately modelling uncertainties (e.g. due to baryonic effects), or incorrectly modelling non-linear effects, or it could be due to new physical signals, e.g. more accurate modelling of non-linear effects introducing new signals that were neglected with other methods.

In this section, we perform an auto-comparison MCMC for each code in which we try to fit a fiducial power spectrum with the same predictor code as the one used to compute the fiducial\textsuperscript{15}. In such a comparison, no significant difference in the best-fit cosmologies between the different codes is expected and differences in the posterior probabilities are limited to different sensitivity contours.

#### 4.1 Sensitivity to the cosmological parameters

We first compare the sensitivity to the cosmological parameters for Euclid galaxy clustering plus (fake) Planck cosmic microwave background (CMB) for each of the three different predictors looked at in this paper. To this end, we analyze four different cases: $\Lambda$CDM and $w_0$CDM as well as two different cut-off scales, $k_{\text{max}} = 0.2\ h\ Mpc^{-1}$ and $k_{\text{max}} = 0.4\ h\ Mpc^{-1}$. The 1D marginalized posterior standard deviations for each parameter in all four cases and for all three predictors are summarized in Table B1. The relative differences of these standard deviations are graphically represented in Fig. 4 for better understanding.

Clearly, this relative difference varies dramatically depending on the considered parameter: for instance, the sensitivity of EuclidEmulator1 to $\omega_{m}$ in the case of a $\Lambda$CDM cosmology cut off at $k_{\text{max}} = 0.2\ h\ Mpc^{-1}$ is almost the same compared to the corresponding sensitivity of HMCode (the relative difference is around 1\%). However, $\sigma_{EE}$ for $h$ in the case of a $w_0$CDM cosmology cut off at $k_{\text{max}} = 0.4\ h\ Mpc^{-1}$ is reduced by $\approx 70\%$ compared to the corresponding $\sigma_{\text{HM}}$ value. In other words, EuclidEmulator1 is roughly $70\%$ more sensitive to $h$ than HMCode. Evidently, the standard de-
viations computed with EuclidEmulator1 are always smaller than those computed with either HMCode or HALOFIT. The more complex the cosmological model and also the higher the cut-off scale \( k_{\text{max}} \), the bigger the effect, i.e. the smaller the uncertainty obtained from a EuclidEmulator1-based fit compared to a HALOFIT- or HMCode-based fit.

In the same Fig. 4, we also show the average values of those relative differences over all parameters (for each studied case). These averaged values range from only 0.9\% (HMCode vs EuclidEmulator1, ΛCDM, \( k_{\text{max}} = 0.2 \text{ h Mpc}^{-1} \)) to 20.4\% (HMCode vs EuclidEmulator1, \( w_0 \)CDM, \( k_{\text{max}} = 0.4 \text{ h Mpc}^{-1} \)) depending on the case. This quantity serves as a proxy for the volume (to the \( 1/d \) power, where \( d \) is the dimension of the parameter space under investigation) of the 1\% credible region of the posterior distribution and thus for the sensitivity to the cosmological parameters. As is stated above and clearly visible in Fig. 4, this average value is always positive, meaning that on average the 1D marginalized posterior standard deviations computed from a EuclidEmulator1-based posterior is smaller than the corresponding value of a posterior based on either HALOFIT or HMCode. It also becomes clear, that the effect is bigger the more complex the cosmological model under consideration and the higher the cut-off scale \( k_{\text{max}} \).

EuclidEmulator1 models the matter power spectrum much more accurately than either of the halo model-based models, particularly for \( 0.1 \text{ h Mpc}^{-1} \leq k \leq 5 \text{ h Mpc}^{-1} \), as is shown in Euclid Collaboration: Knabenhans et al. (2019). However, it is not guaranteed that a more accurate model produces smaller posterior contours (only less biased ones). The fact that we still observe smaller contours here may be explained by the fact that EuclidEmulator1 models the BAOs more accurately than the other two predictors, i.e. scales that carry lots of valuable information about the cosmology. It should, however, not be forgotten that EuclidEmulator1 does not take baryonic effects into account, while HMCode can. But we note that we did not take them into account even for the HMCode runs, instead fixing the nuisance parameters related to baryonic physics to the default values (see Table 1).

Answer to question #1: Although EuclidEmulator1 models the \( P(k) \) more accurately than HALOFIT or HMCode, it also leads, on average, to better sensitivity for the cosmological parameters than the latter two codes, with the largest improvements seen for \( h \) and \( w_0 \) as EuclidEmulator1 is better at modelling the effect of these parameters on the shape of the matter power spectrum.

4.2 The effect of including mildly non-linear scales

In order to establish the importance of mildly non-linear scales in cosmological parameter estimations we perform auto-comparison tests for two different cut-off scales: \( k_{\text{max}} = 0.2 \text{ h Mpc}^{-1} \) and \( k_{\text{max}} = 0.4 \text{ h Mpc}^{-1} \). Although all predictors used in this study could actually reach more strongly non-linear scales, we do not go beyond \( 0.4 \text{ h Mpc}^{-1} \) because we use only a very simple, linear galaxy bias model which is not valid at small scales.

We quantify the effect of including mildly non-linear scales on the posterior sampling by comparing the sizes of the resulting contours. Of course, a more in-depth analysis employing a more sophisticated galaxy bias model needs to be conducted in order to draw conclusion on the effect of including strongly non-linear scales. As this paper is meant to convey a proof of concept, we leave these analyses to future work.

To this end, we interpret again the results presented in Table B1. This time, however, we shall not compare the different prediction models but perform all comparisons for each of the predictors separately. In Fig. 5 we plot the relative differences (again in \( \% \)) of the 1D marginalized posterior standard deviations between the two different cut-off scales for each of the power spectrum prediction models. Positive values imply that \( \sigma_{k_{\text{max}} = 0.4 \text{ h Mpc}^{-1}} < \sigma_{k_{\text{max}} = 0.2 \text{ h Mpc}^{-1}} \). This is the case for almost all parameters for all predictors: the effect of increasing the cut-off scale is largest for EuclidEmulator1 where the posterior standard deviation averaged over all parameters is reduced by 18.2\% in the ΛCDM scenario and by 20.9\% in the \( w_0 \)CDM case. The minimal effect found is 9.2\% (ΛCDM) and 12.1\% (\( w_0 \)CDM), respectively. As above, we have represented the average effect (over all parameters) as solid horizontal lines.

Answer to question #2: The uncertainties can be significantly reduced by increasing \( k_{\text{max}} \) from \( 0.2 \text{ h Mpc}^{-1} \) to \( 0.4 \text{ h Mpc}^{-1} \). This is not only true for EuclidEmulator1 (though for it the effect is largest) but for all tested prediction models. The magnitude of the effect depends strongly on the model and on the predictor under consideration. This observation does not come as a surprise: it is expected that there is valuable information stored in the galaxy-galaxy power spectrum on (mildly) non-linear scales that can be leveraged in order to estimate the cosmological parameters. The effect is expected to grow as the cut-off scale increases. We reiterate, however, that increasing \( k_{\text{max}} \) beyond \( 

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5 MOCK PARAMETER ESTIMATION

Usually, for forecasting we use the same (non-linear) prediction model for cosmological parameter estimations as is used to compute the fiducial cosmology, i.e. auto-comparison tests are the common choice (like in the previous section). This is fine as long as we are interested in the sensitivity of the model response to the choice of parameters (Perotto et al. 2006). We can, however, mimic the situation of data analysis by e.g. taking EuclidEmulator1 to be the model that produces the “observed” data which we try to fit with different predictors, which we refer to as the “cross-comparison” case. I.e., in this section, we always produce the “observed” mock data with EuclidEmulator1. Running such cross-comparison tests using different predictors for the fitting are likely to lead to different best-fit cosmologies (as we will also confirm later), as different codes encode different physics and differ in the way they are implemented. Such tests can hence be used to emphasize that any cosmological code different physics and differ in the way they are implemented.

Such tests can hence be used to emphasize that any cosmological parameter estimation result significantly depends on the code used, unless proper care is taken to account for uncertainties on non-linear scales.

5.1 Code dependence of best-fit cosmology

The third of our key questions (see section 3) is concerned with the choice of the power spectrum prediction code and its effect on the best-fit cosmology. In this subsection, we analyze and interpret the results from cross-comparison tests, where we compare power spectrum fits computed with HALOFIT and HMCode with a fiducial power spectrum computed with EuclidEmulator1. The resulting posteriors for the ΛCDM case are shown in Fig. 6 (left panel: $k_{\text{max}} = 0.2 \, h \, \text{Mpc}^{-1}$; right panel: $k_{\text{max}} = 0.4 \, h \, \text{Mpc}^{-1}$) while the ones for the $w_0$CDM case are shown in Fig. 7 (left panel: $k_{\text{max}} = 0.2 \, h \, \text{Mpc}^{-1}$; right panel: $k_{\text{max}} = 0.4 \, h \, \text{Mpc}^{-1}$). The numerical results for mean posterior cosmologies and the associated 68% credible contours can be found in Table B2.

Looking at the figures showing the cases for $k_{\text{max}} = 0.2 \, h \, \text{Mpc}^{-1}$, we can easily see that for both cosmological models, all codes largely agree on the best-fit cosmology if only $k \leq k_{\text{max}} = 0.2 \, h \, \text{Mpc}^{-1}$ are considered. More concretely, for these cases the fiducial cosmology always lies within the 1σ credible region. The uncertainties are broadly comparable, as only for $h$ (or equivalently $H_0$) is the sensitivity of EuclidEmulator1 considerably better than that of the other two codes. This agreement is not surprising at all, remembering that the three considered power spectrum prediction codes agree almost perfectly on linear scales (Euclid Collaboration: Knabenhans et al. 2019, 2021).

However, increasing the cut-off scale $k_{\text{max}}$ to $0.4 \, h \, \text{Mpc}^{-1}$ paints a different picture (see right panels of Figures 6 and 7): As on these scales the the power spectra predicted by the codes start to deviate more significantly from one another, in this case clear biases between the best-fit cosmologies computed with the three different predictors are seen. The values vary between 1 and 6σ depending on the parameter, but irrespective of the cosmological model. The averaged biases are about 0.5σ if a cut-off scale of $k_{\text{max}} = 0.2 \, h \, \text{Mpc}^{-1}$ is employed and it is $1 - 2\sigma$ if $k_{\text{max}} = 0.4 \, h \, \text{Mpc}^{-1}$. This aspect is visualized in a more condensed way in Fig. 8. In this figure, the distances between the mean cosmology $\mu_{\text{mod}}$ of each model to the fiducial cosmology is given in units of the 1D marginalized standard deviations $\sigma_{\text{mod}}$ for each cosmological model, cut-off scale and cosmological parameter. The RMS distance (averaged over all cosmological parameters) is again indicated by the solid horizontal lines. It is most evident that increasing the cut-off scale from $0.2 \, h \, \text{Mpc}^{-1}$ to $0.4 \, h \, \text{Mpc}^{-1}$ has a major effect on this distance, both on average as well as for some of the individual cosmological parameters. This result is independent of the power spectrum predictor or cosmological model.
Figure 7. $w_0\text{CDM}$ cross comparison of HMCCode and HALOFIT vs EuclidEmulator1 considering $k \leq 0.2\,h\,\text{Mpc}^{-1}$ (left triangle plot) and $k \leq 0.4\,h\,\text{Mpc}^{-1}$ (right triangle plot), respectively. Just as in Fig. 6 we see good agreement between the posteriors in the case with the lower cut-off scale while increasing that scale leads to emergent differences.

Figure 8. Differences between mean ($\mu$) and fiducial cosmology per cosmological parameter for HALOFIT and HMCCode in the cross-comparison setup and for both cut-off scales. The differences are given in units of the respective 1D marginalized posterior standard deviations ($\sigma$). The solid horizontal lines indicate the RMS distances averaged over all parameters in each panel. This plot shows nicely how different implementations of (even mildly) non-linear physics manifest themselves as larger biases in parameter estimation analyses. Moreover, although the RMS distance for HMCCode and HALOFIT happen to be almost identical, the biases in each cosmological parameter can vary drastically (as e.g. in the case of $n_s$).

We trace back the significant biases observed in the right panel plots in figures 6 and 7 to the differences in the power spectrum predictions between the different predictor codes. To do so we first look at the quantities

$$Q_{\text{mod}}(k; p) \equiv \frac{P_{\text{mod}}(k; p)}{P_{\text{EE}}(k; p_{\text{fid}})},$$

and

$$\Delta Q_{\text{mod}}(k; p) \equiv Q_{\text{mod}}(k; p) - Q_{\text{mod}}(k; p_{\text{fid}}),$$

shown in Fig. 9. Here $\text{mod} \in \{\text{HF, HM}\}$, the set of investigated cosmological parameters is $p \in \{\omega_b, \omega_m, n_s, h, w_0, A_s\}$, $p_{\text{fid}}$ corresponds to the respective parameter set to its fiducial value as given in Table 1. The redshift is kept fixed at $z = 1$. The quantity $Q$ differs from quantity $R$ defined in Eq. 1 by keeping the denominator fixed to the fiducial cosmology, while it is also varied in the quantity $R$.

Since the biases grow significantly as we increase the cut-off scale from $k = 0.2\,h\,\text{Mpc}^{-1}$ to $k = 0.4\,h\,\text{Mpc}^{-1}$, we shall focus primarily on this region in the following discussion (marked by the vertical indigo bands in Fig. 9). The argument is most eye-catchingly illustrated considering the example of $n_s$ (third row on Fig. 9): while in general (i.e. over the entire $k$ range taken into account for the MCMCs) the quantities $Q$ are very similar for HALOFIT and for HMCCode, they show a difference in behaviour between $k = 0.2\,h\,\text{Mpc}^{-1}$ and $k = 0.4\,h\,\text{Mpc}^{-1}$. The curves turn upward for HALOFIT but downward for HMCCode. In other words, at the small scale end of the considered $k$ range, HALOFIT tends to overestimate the power spectrum while HMCCode underestimates it. The overestimation for HALOFIT is however alleviated if the $n_s$ value is decreased compared to the fiducial value. Similarly, the underestimation in the case of HMCCode is reduced if $n_s$ is somewhat increased. Hence, the MCMC favours lower $n_s$ values for HALOFIT and higher ones for HMCCode when fitting the fiducial cosmology. This aligns perfectly with what we observe in figures 6 and 7. The same argument can be applied to all other cosmological parameters, albeit with complex parameter degeneracies (made more complex by the inclusion of mock Planck data) resulting in shifts for most parameters.
Figure 9. This figure shows the change of the power spectrum predicted by either HALOFT or HMCode as individual cosmological parameters are varied while all other parameters are fixed to the fiducial cosmology. The comparison is shown in terms of the quantity $Q_{\text{model}}$, being defined in the main text as the ratio of a power spectrum predicted by model $\in \{\text{Halofit, HMCode}\}$ with respect to the power spectrum produced by EuclidEmulator1 at the fiducial cosmology. The curves are shown plotted in the full range $k \in [0.005, 5]$, however, only the colored parts are actually taken into account in the MCMCs and hence only those $k$ modes have an effect on the biases observed in the MCMC posteriors in this section. The shaded indigo bands mark the region between $k = 0.2 h \, \text{Mpc}^{-1}$ and $k = 0.4 h \, \text{Mpc}^{-1}$ which is responsible for the larger biases observed in the analysis with the larger cut-off scale. Notice the differences and similarities between this figure and Fig. 2.

E.g., changing $\omega_m$ (second row on Fig. 9) helps adjust for the differences on BAO scales for HMCode (note the posterior for $\omega_m$ remains largely unchanged for HALOFT compared to EuclidEmulator1), while also modulating the small scale amplitude shift from changing $n_s$ with a similar but somewhat different scale dependence: where increasing $n_s$ increases power on small scales and decreases power on large scales, instead decreasing $\omega_m$ decreases power across the observable range, except for the very largest scales where the impact is negligible, with a larger effect on smaller scales.

Therefore, the combination of increased $n_s$ and decreased $\omega_m$ that we see for HMCode results in a different impact on the matter power spectrum than changing $n_s$ alone, giving a flatter increase in power from large to smaller scales.

Also, we see only a very small shift in the mean of each posterior for $H_0$, but instead see inflated posteriors. We can understand this as the mock Planck data not allowing for a considerably different mean value from the fiducial value, but that the sampler attempts to adjust for differences on BAO scales by varying $H_0$, resulting in a decrease in sensitivity to $H_0$ compared to the EuclidEmulator1 case, where the non-linear prescription matches the fiducial model, due to washing out information from the placement of the BAO peaks.
Finally, we note that $\omega_b$ is largely unaffected by biases, as this parameter is very well determined by the mock Planck likelihood and cannot easily be shifted.

**Answer to question #3:** While the choice of the power spectrum predictor only has a relatively minor impact on the best-fit cosmology and its uncertainties when mostly linear scales are considered, the impact grows as increasingly larger $k$ modes are taken into account for the fit. We found this result independently of the fact whether a simple $\Lambda$CDM or a more complex $w_0$CDM model was considered.

### 5.2 Effect of theoretical uncertainty models on the posterior distribution

In the previous subsection, we have shown that the choice of the power spectrum predictor can, in some cases, have a quite significant effect on the resulting best-fit cosmology. These differences clearly cannot be physical but must be due to modelling and implementation differences. In order to account for this, as was discussed above, one has to include yet another source of uncertainty into the modelling: a theoretical uncertainty (see section 2.4). However, a priori such an uncertainty is not known and hence it has to be modelled.

For this reason it is a valid question to ask what the impact of the choice for a theoretical uncertainty model is. To answer this question we ran two sets of MCMCs for EuclidEmulator1, HALOFIT and HMCde (restricting ourselves to the $w_0$CDM model and the case where $k_{\text{max}} = 0.4 \, h \, \text{Mpc}^{-1}$): one set with the theoretical uncertainty envelope as given in Sprenger et al. (2019) referred to as $\sigma_{\text{th,KB}}$ and one set with the envelope described in section 2.4 referred to as $\sigma_{\text{th,S}}$. More concretely, we first compute a fiducial galaxy power spectrum $P_{gg}$ based on EuclidEmulator1. In a second step, we run MCMCs for all three predictors, always comparing to the same fiducial. This means that in the case of EuclidEmulator1 we...
perform an auto-comparison test, while for HALOFIT and HMCode cross-comparisons are performed.

The different choices of the theoretical uncertainty manifest themselves in different distances between the best-fit cosmologies and in different contours. Since $\sigma_{th, KB}$ was modelled with specific information about the codes being used in this analysis while $\sigma_{th, S}$ is a rather generic envelope, we expect the posterior distributions featuring $\sigma_{th, KB}$ to agree significantly better with each other than the ones computed with $\sigma_{th, S}$. Specifically, the more aggressive theoretical uncertainty $\sigma_{th, KB}$ down-weights the non-linear portion and BAO scales where the codes disagree relative to the linear portion of the galaxy power spectrum where the codes agree very well, so the bias is reduced. While this test mostly serves as a sanity check, it helps us understand the often neglected impact of modelling choices on the final cosmological parameter estimation.

At this point we shall reiterate: It is always possible to implement so much theoretical uncertainty into an MCMC such that any two codes would agree with each other. Hence, while there is no real challenge in making the best-fit cosmologies of MCMCs based on EuclidEmulator1, HALOFIT and HMCode agree with each other, the difficulty lies in doing so without significantly inflating the credibility contours.

Our test results are shown in Fig. 10 and more compactly in Fig. 11 (for numerical results see Table B3). The filled contours in 10 represent the MCMCs based on $\sigma_{th, KB}$ and the empty contours those based on $\sigma_{th, S}$. The empty contours correspond to the contours shown in Fig. 7. It is evident that the filled contours agree significantly better with each other than the empty ones while their areas are not much larger. The upper panels of Fig. 11 show again the RMS distances (biases) between the mean cosmologies in units of the 68% 1D marginalized posterior standard deviations $\sigma$ (similar to Fig. 8). Clearly, the biases are smaller over all parameters when $\sigma_{th, KB}$ is used compared to employing $\sigma_{th, S}$. However, in order to paint the complete picture we need to take into account that not only the biases but also the uncertainties (i.e. the $\sigma$’s themselves) change when different uncertainty envelopes are used. The ratio between the standard deviations $\sigma$ depending on the uncertainty envelope $\sigma_{th,i}$, $i \in \{S, KB\}$ is shown in the lower panel. As expected, this ratio is always greater or equal to unity implying that the uncertainties are somewhat increased when $\sigma_{th, KB}$ is used compared to using $\sigma_{th, S}$. Note, however, how the maximal ratio is only $\sim 1.4$. This shows that a carefully modelled uncertainty envelope is capable of significantly reducing biases while not inflating the uncertainty contours by a lot.

Answer to question #4: The choice of the theoretical uncertainty envelope $\sigma_{th}$ can have a significant impact on the resulting estimate of the best-fit cosmological parameter values. A theoretical uncertainty which is not modelled with enough care may lead to a considerable bias in the best-fit cosmology. On the other hand, a carefully modelled theoretical uncertainty is able to account for model specifics and make the resulting posterior less dependent on the chosen power spectrum predictor while keeping the posterior credibility contours reasonably tight. We hence strongly suggest that in future cosmological parameter estimation studies more focus shall be put on the aspect of theoretical uncertainties in order to make the results more model-independent.

6 CONCLUSION

In this paper, we have studied cosmological parameter inference results based on galaxy clustering observables in both $\Lambda$CDM and $w_0$CDM cosmologies focusing on the following four major questions:

(i) What are the performance differences of EuclidEmulator1, HALOFIT and HMCode?
(ii) What is the added value of considering (mildly) nonlinear scales in parameter forecasts?
(iii) What is the impact of the choice of the (non-linear) predictor model on the parameter estimation result?
(iv) How are the forecasting results affected by different choices of theoretical uncertainty models?

To answer these questions, we have run multiple MCMCs with the code MontePython. The three different power spectrum predictors used for this work are EuclidEmulator1, HALOFIT and HMCode. Notice that EuclidEmulator1 allows to predict the nonlinear correction only. We take advantage of this fact by modifying the Kaiser formula as described in appendix A.

We start out our analysis (see section 4.1) by a simple comparison of the MCMC posteriors based on the three different power spectrum predictors. To do so, we let MontePython estimate the parameters of a fiducial model, where both the fits and the fiducial are computed with the same predictor (referred to as “auto-comparison” tests). It is therefore expected that the best-fit cosmology recovers the fiducial cosmology almost perfectly in all cases. However, the
posterior contours of EuclidEmulator1 tend to be tighter than the contours of HALOFIT and HMCode. The magnitude of this effect varies significantly depending on the cosmological model, the non-linear prescription and the $k$-mode cut-off employed. The minimal value is 2.2% in the case of HMCode vs. EuclidEmulator1 for $\Lambda$CDM with $k_{\text{max}} = 0.2 \, h \, \text{Mpc}^{-1}$ while the maximum value is 27.9% the case of HALOFIT vs. EuclidEmulator1 for $w_0$CDM with $k_{\text{max}} = 0.4 \, h \, \text{Mpc}^{-1}$. These results are shown in Fig. 4. This observation together with the knowledge that the physics of non-linear clustering is more accurately modelled in EuclidEmulator1 than in the halo model-based models points at the fact that parameter forecasts based on EuclidEmulator1 are subject to smaller uncertainties than those based on the other two predictors.

We analyze the effect of including (mildly) non-linear scales on the posterior probability distribution of the MCMCs in section 4.2. Concretely, we run two sets of MCMCs, one considering only modes $k \leq k_{\text{max}} = 0.2 \, h \, \text{Mpc}^{-1}$ (mostly linear scales) and one including also modes $k \leq k_{\text{max}} = 0.4 \, h \, \text{Mpc}^{-1}$ (linear and mildly non-linear scales). We find that doing so reduces the 1D marginalized posterior standard deviations per parameter for all three different predictors (see Fig. 5). This finding confirms that there is valuable information at smaller spatial scales that can be leveraged in order to learn more about the cosmological parameters (clearly this is the motivation for many modern cosmological surveys such as Euclid to investigate data from those smaller scales). The magnitude of the effect varies between $\sim 9\%$ and $\sim 25\%$ (on average over all tested cosmological parameters) depending on the predictor and the cosmological model under consideration. Notice that we did not include scales above $0.4 \, h \, \text{Mpc}^{-1}$ as we used only a simple linear galaxy bias model and we neglected baryonic effects in our analysis. The validity of such a simple model is known to break at even more non-linear scales. We thus consider the results presented in this paper as a proof of concept and leave a more sophisticated analysis of the impact of strongly non-linear scales on the posterior uncertainties to future work. Nevertheless, from these observations we can confirm that extending the range of considered $k$ modes beyond linear scales is certainly beneficial to a relevant degree for future cosmological parameter forecasts. In addition we find that, when more non-linear scales are considered in such an analysis, then the results depend on the prescription used to model the physics at these scales, unless a proper theoretical uncertainty is used.

In section 5, we go one step further and take on the view point of a mock data analysis than merely a parameter forecast. In a data analysis, we need to consider the possibility that the best-fit cosmology is biased compared to the true cosmology underlying the data. We test two different theoretical uncertainty envelopes in section 5.2: a rather generic one and a carefully modelled one which is actually informed by the agreement of the three power spectrum predictors compared with each other. We find that the latter envelope leads to much smaller biases in the best-fit cosmologies while only marginally inflating the credible contours. By using our new theoretical uncertainty $\sigma_{\text{th},KB}$, the RMS distance between the mean and the fiducial cosmology can be reduced by a factor of $\sim 2$ for HALOFIT and even by a factor of $\sim 5$ for HMCode. This result is visualized in figures 11 and 10. We thus advocate for the consideration of carefully modelled theoretical uncertainty envelopes in future data analyses as we demonstrate that they render the results much less model-dependent.

We conclude by stating that we have found the tightest posterior with the smallest bias between the best-fit and the fiducial cosmology by taking into account not just linear but also mildly non-linear scales and by including a well-informed theoretical uncertainty envelope into our model.

GLOSSARY

Codes:

CLASS

Cosmological linear anisotropy solving system

EuclidEmulator1

Emulator code to emulate non-linear corrections (boost factors) to DM power spectra 1–5, 7–21

EuclidEmulator2

Successor of EuclidEmulator1. Emulator code to emulate non-linear corrections (boost factors) to DM power spectra self-consistently including corrections from massive neutrinos and dark energy perturbations.

HALOFIT

Analytical code to produce non-linear power spectra 1–3, 5, 7–21

HMCode

Analytical code to produce non-linear power spectra 1–5, 7–21

MontePython

Monte Carlo code for Cosmological Parameter extraction. It contains likelihood codes of most recent experiments, and interfaces with the Boltzmann code class for computing the cosmological observables.

Acronyms:

BAO

Baryonic acoustic oscillations

CMB

Cosmic microwave background

DE

Dark energy

EoS

equation of state

MCMC

Markov chain Monte Carlo

NLC

nonlinear correction

PT

perturbation theory
RSD
redshift space distortion 6, 8, 17

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DATA AVAILABILITY

The data underlying this article will be shared on reasonable request to the corresponding authors.

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APPENDIX A: QUASI-NON-LINEAR GALAXY-GALAXY POWER SPECTRUM

In section 3 of the main text we introduced how the NLC computed by EuclidEmulator1 can be used to derive the observable galaxy-galaxy power spectrum. However, we glossed over a fact that the relation we used is not fully self-consistent. This aspect is analyzed in more depth in this appendix.

A1 Review of linear RSDs modelling

Eq. 7 explains how the dark matter power spectrum $P_{\delta\delta}(k,z)$ is related to the observable galaxy-galaxy power spectrum $P_{gg}(k,\mu,z)$. It is repeated here for convenience:

$$P_{gg}(k,\mu,z) = \frac{f_{AP}(z)}{f_{RSD}(k,\mu,z)} \times f_{res}(k,\mu,z) \times f_{\delta\delta}(\hat{k},\hat{\mu},z) \times b^2(z) \times P_{\delta\delta}(k,z),$$

(A1)
Further, in the same section we showed how the the NLC \( B \) computed by EuclidEmulator1 can be incorporated into that equation. In this section, we discuss the relation between \( P_{gg} \) and \( B(k,z) \) in more detail, thereby focusing on the RSDs (\( \delta_{RSD} \)) and ignore the galaxy bias \( (b) \), Alcock-Paczynski (\( f_{AP} \)), and resolution (\( f_{res} \)) correction factors.

The observable galaxy-galaxy power spectrum depends on the direction of observation (see Eq. A1) because in addition to the redshift signal due to the Hubble expansion there is a contribution to the redshift distance caused by the peculiar velocities of the galaxies, i.e.

\[
s = cz = H_0 d + v_r
\]

where \( s \) denotes the redshift distance to a galaxy, \( c \) is the speed of light, \( z \) is the measured redshift of a galaxy, \( H_0 \) the Hubble parameter at present day, \( d \) is the proper distance to the galaxy and \( v_r \) is the peculiar radial velocity. On large, linear scales the RSD effect is at present day, light, \( s \), the dark matter velocity power spectrum \( P_{\delta \delta} \), the dark matter mass-velocity cross-power spectrum \( P_{\delta \theta} \), the dark matter power spectrum \( P_{\theta \theta} \), and the dark matter mass-velocity cross-power spectrum \( P_{\delta \theta} \).

\[
P_{gg}(k,\mu,z) = P_{\delta \delta}(k,z) + 2\mu^2 \beta \bar{P}_{\delta \theta}(k,z) + \mu^4 \bar{P}_{\theta \theta}(k,z),
\]

where \( \beta := \frac{f}{b(z)} \) is the ratio of the growth rate \( f \) and the galaxy bias \( b \). Although Eq. A3 holds in the non-linear regime, in practice the computation of \( \bar{P}_{non-linear} \) with this equation requires some means to compute \( \bar{P}_{non-linear} \) and \( \bar{P}_{linear} \). Unfortunately, EuclidEmulator1, HALOFIT and HMCCode do only provide the capability to compute \( \bar{P}_{non-linear} \). However, to linear order the divergence of the velocity field is equal to the mass density field:

\[
\delta(k) \approx \frac{k\delta(x)}{aH} = \theta(k)
\]

such that \( \bar{P}_{linear} = \frac{P_{linear}}{\delta \delta} = \bar{P}_{linear} \). It is thus possible to simplify Eq. A3 to

\[
\bar{P}_{gg}(k,\mu,z) = \bar{P}_{linear}(k,z) + 2\mu^2 \beta \bar{P}_{linear}(k,z) + \mu^4 \bar{P}_{linear}(k,z)
\]

On smaller scales, the Fingers of God effect (Tully & Fisher 1978; Jackson 1972) adds to the RSD and is accounted for by an additional prefactor (Bull et al. 2015) such that in total

\[
\bar{P}_{gg}(k,\mu,z) = f_{RSD} \bar{P}_{linear}(k,z)
\]

with

\[
f_{RSD} = \left[ 1 + \mu^2 f \right]^2 \exp \left( -k^2 \sigma_c^2 \right).
\]

A2 Quasi-non-linear extension of the RSDs

In the previous subsection we summarized the well-known Kaiser formula that describes the RSDs of the linear power spectrum. Because we are interested in analyses including (mildly) non-linear scales in this paper, a generalization of that Kaiser formula to the non-linear galaxy-galaxy power spectrum is required. The obvious generalization is (compare to Eq. A6)

\[
\bar{P}_{gg}^{QNL1}(k,\mu,z) = f_{RSD} \bar{P}_{non-linear}(k,z).
\]

where the superscript “QNL1” stands for “quasi-non-linear”. Notice however that this generalization breaks self-consistency as the relation Eq. A4 cannot be generalized to the non-linear power spectrum. Let’s expand the non-linear matter power spectrum into its linear contribution and the NLC:

\[
f_{RSD}^{QNL1}(k,\mu,z) = f_{RSD} \bar{P}_{linear}(k,z).B(k,z).
\]

Here, we started with Eq. A3, linearized this equation using Eq. A4, factored out \( \bar{P}_{\delta \delta} \) in Eq. A5 and replaced it by its non-linear counterpart. We shall refer to this as the “QNL1” model.

A second, self-consistent option to generalize Eq. A6 is to keep the linear versions of the second and third term in Eq. A3, i.e. keep \( \bar{P}_{linear} = \bar{P}_{linear} = \bar{P}_{linear} \) and only use the non-linear version of \( \bar{P}_{\delta \delta} \) in the first term of Eq. A3. This model, referred to as “QNL2”, is thus given by:

\[
\bar{P}_{gg}^{QNL2}(k,\mu,z) = \bar{P}_{non-linear}(k,z) + 2\mu^2 f \bar{P}_{linear}(k,z) + \mu^4 f^2 \bar{P}_{linear}(k,z)
\]

Notice that self-consistency is not broken at the price of ignoring the non-linear contributions in \( \bar{P}_{\delta \delta} \) and \( \bar{P}_{\theta \theta} \). If we substitute in the NLC the linear matter power spectrum can be factored out again:

\[
\bar{P}_{gg}^{QNL2}(k,\mu,z) = \left[ B(k,z) + 2\mu^2 f + \mu^4 f^2 \right] \bar{P}_{linear}(k,z)
\]

Adding the Fingers of God for RSDs at small scales we arrive at

\[
f_{RSD}^{QNL2}(k,\mu,z) = f_{RSD} \bar{P}_{linear}(k,z)
\]

with the following expression for the RSD correction factor:

\[
f_{RSD}^{QNL2} = \left[ B(k,z) + 2\mu^2 f + \mu^4 f^2 \right] \exp \left( -k^2 \sigma_c^2 \right).
\]

Having two options to generalize the Kaiser formula, the natural question arises by how much \( \bar{P}_{gg}^{QNL1} \) and \( \bar{P}_{gg}^{QNL2} \) differ from one another. Results of a test investigating this question are shown in Fig. A1. As can be seen from these results, for the purpose of parameter forecasting there is virtually no difference between the QNL1 and QNL2 model. For this reason we treat the two models as one for the purposes described in the remaining analyses. More concretely, we have used model QNL2 in combination with EuclidEmulator1 because it is slightly more self-consistent than model QNL1. The latter, however, is simpler to implement when using HALOFIT or HMCCode. Whenever we worked with the latter two codes, we hence applied QNL1, including for the fiducial model in the cross-comparison results of section 5.
Parameter inference with non-linear galaxy clustering

Figure A1. QNL1 vs QNL2 for \( \Lambda \)CDM (top panel) and \( w_0 \)CDM (bottom panel). From these plots it becomes obvious that the two different QNL models (see Eq. 8 and 9) do only very marginally affect the posterior contours. The differences are so small that we neglect them throughout this entire paper.

APPENDIX B: MCMC RESULTS

In this appendix we provide the numerical results of the MCMCs (see tables) as well as the contours plots of the posterior probability distributions of the cosmological parameters as obtained in auto-comparison experiments (i.e. the sensitivity forecasting-like setup discussed in section 4).

B1 Auto-comparison results

In section 4 we described the impact of the choice of the power spectrum predictor on forecasting results. As explained, in this setting only the various sensitivities are of interest as the mean posterior cosmologies coincide with the fiducial cosmology almost perfectly by construction. In the main text we provided the sensitivity information in the form of bar plots (see Fig. 4 and 5). As this representation may be somewhat unusual, we provide the more common contour plots here in this appendix (Fig. B1 and B2).

B2 Cross-comparison results

We have discussed the impact of the choice of the power spectrum predictor on the final parameter estimation results in section 5. There we have provided the graphical summaries of the “cross-comparison” tests. Here, however, we provide the underlying numerical results in tabular form in order to assert full transparency. Table B2 lists the mean cosmology and the 68% credibility limits per cosmological parameter and cosmological model for both power spectrum predictors, HALOFIT and HMCode, with the fiducial cosmology given by EuclidEmulator1. For all results in that table the uncertainty envelope by Sprenger et al. (2019) was used. In Table B3, on the other hand, we show the same quantities for the test with our new uncertainty envelope employed, that we have shown as lime-green lines in Fig. 3. Notice that for this experiment the results shown for EuclidEmulator1 correspond to an auto-comparison test while those for HALOFIT and for HMCode are cross-comparison tests.

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Table B1: Table of 1D marginalized posterior standard deviations ($\sigma$) for auto-comparisons. Each column reports about a power spectrum predictor that was used to both create the fiducial model and to sample the posterior probability surface. In each row of the table the standard deviation of such a posterior along a specific cosmological parameter axis and for a specific cosmological model and cut-off scale is denoted. Comparing the different predictors with each other it turns out that the posterior standard deviations of EuclidEmulator1 are usually smaller than those of HALOFIT or HMCode. The exception is $\omega_m$ (and in some cases $n_s$) which is itself an input parameter to EuclidEmulator1 while in HALOFIT and HMCode it is derived from $\omega_b$ and $\omega_{cdm}$. While in general the size of the posterior contours seem to be comparable for simple ΛCDM cosmologies and low cut-off scales, it can be seen that the difference in $\sigma$ between EuclidEmulator1 and HALOFIT and HMCode is larger for the case of $\omega_0$CDM cosmologies and the larger cut-off scale. The relative difference is sometimes as large as 70% along a single parameter as in the case of the Hubble parameter $h$ for $\omega_0$CDM and $k_{\text{max}} = 0.4$ h Mpc$^{-1}$ (more clearly visible in the lower right panel of Fig. 4). From this one can conclude that EuclidEmulator1 has a substantial positive effect on the confidence of a parameter estimation compared to the other two predictors.

|         | EuclidEmulator (EE) | HALOFIT (HF) | HMCode (HM) |
|---------|---------------------|--------------|--------------|
| $\sigma(\omega_b)$ | $1.11e-04$ | $1.16e-04$ | $1.15e-04$ |
| $\sigma(\omega_m)$ | $4.49e-04$ | $4.22e-04$ | $4.24e-04$ |
| $\sigma(n_s)$ | $2.16e-03$ | $2.34e-03$ | $2.15e-03$ |
| $\sigma(h)$ | $1.49e-03$ | $1.64e-03$ | $1.62e-03$ |
| $\sigma(\sigma_8)$ | $2.79e-03$ | $2.83e-03$ | $2.74e-03$ |
| $\sigma(\omega_b)$ | $1.04e-04$ | $1.16e-04$ | $1.16e-04$ |
| $\sigma(\omega_m)$ | $4.18e-04$ | $4.09e-04$ | $4.11e-04$ |
| $\sigma(n_s)$ | $1.83e-03$ | $1.89e-03$ | $1.79e-03$ |
| $\sigma(h)$ | $1.20e-03$ | $1.53e-03$ | $1.58e-03$ |
| $\sigma(\sigma_8)$ | $2.07e-03$ | $2.18e-03$ | $2.28e-03$ |
| $\sigma(\omega_b)$ | $1.12e-04$ | $1.18e-04$ | $1.15e-04$ |
| $\sigma(\omega_m)$ | $4.74e-04$ | $4.37e-04$ | $4.48e-04$ |
| $\sigma(n_s)$ | $2.26e-03$ | $2.35e-03$ | $2.15e-03$ |
| $\sigma(h)$ | $2.24e-03$ | $3.02e-03$ | $3.02e-03$ |
| $\sigma(\omega_0)$ | $9.36e-03$ | $1.15e-02$ | $1.17e-02$ |
| $\sigma(\sigma_8)$ | $3.23e-03$ | $3.68e-03$ | $3.58e-03$ |
| $\sigma(\omega_b)$ | $1.11e-04$ | $1.14e-04$ | $1.17e-04$ |
| $\sigma(\omega_m)$ | $4.37e-04$ | $4.17e-04$ | $4.26e-04$ |
| $\sigma(n_s)$ | $1.95e-03$ | $1.91e-03$ | $1.82e-03$ |
| $\sigma(h)$ | $1.62e-03$ | $2.74e-03$ | $2.77e-03$ |
| $\sigma(\omega_0)$ | $7.56e-03$ | $9.97e-03$ | $1.02e-02$ |
| $\sigma(\sigma_8)$ | $2.34e-03$ | $2.74e-03$ | $2.81e-03$ |
Figure B1. Auto-comparisons for $\Lambda$CDM (top) and $w_0$CDM (bottom) with cut-off scale $k_{\text{max}} = 0.2 \, h \, \text{Mpc}^{-1}$.

Figure B2. Auto-comparisons for $\Lambda$CDM (top) and $w_0$CDM (bottom) with cut-off scale $k_{\text{max}} = 0.4 \, h \, \text{Mpc}^{-1}$.
Table B2: Mean posterior cosmologies and 68% credible intervals for HALOFIT and HMCode as resulting from cross-comparison MCMCs where EuclidEmulator data serves as mock data. The fiducial values for the cosmological parameters are given in the first data column. Notice that for all results presented in this table the theoretical uncertainty by Sprenger et al. (2019) is used (for results based on the new theoretical uncertainty introduced in this paper see Table B3).

|                | Fiducial          | Halofit (HF)          | HMCode (HM)          |
|----------------|-------------------|-----------------------|----------------------|
| **ACDM**       |                   |                       |                      |
| 0.2 \(h \text{ Mpc}^{-1}\) | \(\omega_b\) 0.02200 | 0.02193 ± 0.00012     | 0.02194 ± 0.00011    |
|                | \(\omega_m\) 0.1432 | 0.1431 ± 0.00044      | 0.1428 ± 0.00044     |
|                | \(n_s\) 0.9600   | 0.9593 ± 0.00237      | 0.9577 ± 0.00217     |
|                | \(h\) 0.6700     | 0.6696 ± 0.00170      | 0.6705 ± 0.00168     |
|                | \(\sigma_8\) 0.8300 | 0.8287 ± 0.00278      | 0.8278 ± 0.00270     |
| **w_0CDM**     |                   |                       |                      |
| 0.2 \(h \text{ Mpc}^{-1}\) | \(\omega_b\) 0.02200 | 0.02193 ± 0.00012     | 0.02193 ± 0.00012    |
|                | \(\omega_m\) 0.1432 | 0.1433 ± 0.00040      | 0.1418 ± 0.00043     |
|                | \(n_s\) 0.9600   | 0.9535 ± 0.00194      | 0.9715 ± 0.00184     |
|                | \(h\) 0.6700     | 0.6676 ± 0.00158      | 0.6753 ± 0.00161     |
|                | \(\sigma_8\) 0.8300 | 0.8223 ± 0.00202      | 0.8405 ± 0.00252     |
| 0.4 \(h \text{ Mpc}^{-1}\) | \(\omega_b\) 0.02200 | 0.02193 ± 0.00011     | 0.02197 ± 0.00012    |
|                | \(\omega_m\) 0.1432 | 0.1433 ± 0.00042      | 0.1415 ± 0.00046     |
|                | \(n_s\) 0.9600   | 0.9535 ± 0.00194      | 0.9715 ± 0.00186     |
|                | \(h\) 0.6700     | 0.6669 ± 0.00274      | 0.6684 ± 0.00288     |
|                | \(w_0\) -1.0000  | -0.9969 ± 0.00974     | -0.9695 ± 0.01065    |
|                | \(\sigma_8\) 0.8300 | 0.8216 ± 0.00272      | 0.8367 ± 0.00292     |

Table B3: Mean posterior cosmologies and 68% credible intervals for HALOFIT and HMCode as resulting from cross-comparison MCMCs where EuclidEmulator data serves as mock data. The fiducial values for the cosmological parameters are given in the first data column. All results presented in this table are based on the new theoretical uncertainty given in section 2.4.2.

|                | EuclidEmulator (EE) | Halofit (HF) | HMCode (HM) |
|----------------|---------------------|--------------|--------------|
| **w_0CDM**     |                    |              |              |
| 0.4 \(h \text{ Mpc}^{-1}\) | \(\omega_b\) 0.02200 ± 0.00011 | 0.02202 ± 0.00012 | 0.02199 ± 0.00012 |
|                | \(\omega_m\) 0.1432 ± 0.00050 | 0.1430 ± 0.00048 | 0.1424 ± 0.00052 |
|                | \(n_s\) 0.9600 ± 0.00234 | 0.9562 ± 0.00240 | 0.9651 ± 0.00235 |
|                | \(h\) 0.6700 ± 0.00280 | 0.6697 ± 0.00385 | 0.6674 ± 0.00394 |
|                | \(w_0\) -0.9997 ± 0.01090 | -0.9977 ± 0.01324 | -0.9794 ± 0.01355 |
|                | \(\sigma_8\) 0.8300 ± 0.00328 | 0.8246 ± 0.00365 | 0.8293 ± 0.00375 |