A wealth distribution model on isolated discrete time domains, which allows the wealth to exchange at irregular time intervals, is used to describe the effect of agent’s trading behavior on wealth distribution. We assume that the agents have different degrees of risk aversion. The hyperbolic absolute risk aversion (HARA) utility function is employed to describe the degrees of risk aversion of agents, including decreasing relative risk aversion (DRRA), increasing relative risk aversion (IRRA), and constant relative risk aversion (CRRA). The effect of agent’s expectation on wealth distribution is taken into account in our wealth distribution model, in which the agents are allowed to adopt certain trading strategies to maximize their utility and improve their wealth status. The Euler equation and transversality condition for the model on isolated discrete time domains are given to prove the existence of the optimal solution of the model. The optimal solution of the wealth distribution model is obtained by using the method of solving the rational expectation model on isolated discrete time domains. A numerical example is given to highlight the advantages of the wealth distribution model.

1. Introduction

In recent years, unequal distribution of wealth has caused a series of negative phenomena in the society. Empirical and theoretical research on wealth distribution has become an active branch of modern economics. The problem of wealth distribution is closely related to the optimal conditions of economic growth. Studying the distribution of wealth is important for analyzing the wealth inequality and formulating tax and redistributive policies.

The dynamic distribution of wealth is a complex phenomenon, which is determined by many factors, including tax, welfare, innovation, education, consumption, and agents’ preferences. Guala [1] studies the redistribution of wealth among agents in the presence of taxation in the trading process. Pareschi and Toscani [2] consider the influence of agent’s knowledge level on wealth distribution and analyze the formation process of knowledge. Toscani et al. [3] discuss the influence of agents’ preferences on wealth distribution and get Pareto curve of wealth distribution. Brugna and Toscani [4] generalize the works of Toscani et al. [3] and further investigate the changes in wealth distribution of two types of agents with different trading strategies. A. S. Chakrabarti and B. K. Chakrabarti [5] propose a model based on consumers’ optimization which can give rise to those particular forms of asset exchange equations and discuss a problem of income distribution. In [6], the authors further investigate income and wealth inequality in the context of a generalized kinetic exchange model and conclude that the economy indeed shows inequality reversal under some certain conditions. Chakrabarti [7] discusses the laws of wealth distribution, firm size distribution, and the city size distribution in a common framework and shows that the equilibrium configurations of some general economic mechanisms are consistent with a power law in general and Zipf’s law in particular, in size distribution. A survey of wealth distribution is referred to [8].

In fact, wealth distribution is not only related to the knowledge level and preference of agents but also to the transaction time. Therefore, it is necessary to divide the agent’s transaction into several periods and set up a discrete wealth distribution model. Chatterjee et al. [9] assume that
the agent’s saving tendency satisfies a certain distribution and give a discrete wealth distribution model. Benhabib et al. [10] investigate the wealth distribution in Bewley economies with capital income risk and solve an infinite horizon consumption-saving problem with incomplete markets. In [10], the equilibrium distribution for consumption, savings, and wealth is obtained under the assumption that agents face saving restrictions. Mariacristina and Giulio [11] study the influence of agent’s savings on wealth distribution and find out the factors causing wealth inequality. Other discrete models on wealth distribution can be found in [12–16].

In actual transactions, agents usually anticipate the price of assets or commodities in the market based on past trading information. This anticipated behavior affects the distribution of wealth among agents. Aceldonksi [17] discusses the influence of agent’s expectation on wealth distribution. The results show that the effect of agent’s expectation on wealth distribution is obvious. In [17], a standard constant relative risk aversion (CRRA) utility function is used to describe the behaviors of agents, and a discrete wealth distribution model affected by agent’s expectation is given. In [10, 17], although the agent’s transaction is divided into several periods, the time interval of the transaction is uniform, without considering the wealth distribution when the transaction interval is uneven.

Generally speaking, trading is not a uniformly discrete nor continuous behavior. Consider the stock market as an example. In a stock market, an agent buys shares of a particular company in varying time intervals. Suppose that on the first week of the year, the agent buys stocks every day, and in the next several weeks, the company generates a good profit, which makes the agent buy stocks several times a week. Unfortunately, due to the company’s poor performance and lower earnings for the rest of the year, the agent buys only one stock on the last week of the year. Since there is no rule to restrict the agent’s trading behavior, we do not know whether the time interval of trading is uniform. As a result, the agent’s trading behavior needs to be described in isolated discrete time domains where events may occur at unevenly spaced time points. Atici et al. [18] present a dynamic optimal problem from economics and construct a time scale model. Atici et al. [19] investigate a perfect-foresight utility maximisation problem by using a method for solving the dynamic optimisation problem in which the real-valued function and constraints depend on different times. Atici and Turhan [20] discuss the deterministic dynamic sequence problem on isolated discrete time domains. Atici et al. [21] introduce a nonlinear stochastic dynamic problem on isolated discrete time domains and give the Euler equation and transversality condition for the problem. The work in [21] provides a method for solving the nonlinear growth model on isolated discrete time domains. Other studies about the isolated discrete models can be found in [22–25].

In [10, 17, 21], a CRRA utility function is utilized to describe the agent’s behavior. In actual transactions, the agent’s aversion for risk is complex. To better study the distribution of wealth among agents, we employ the hyperbolic absolute risk aversion (HARA) utility function to characterize the degrees of risk aversion of agents. The HARA utility function includes several agents’ aversion for risk, such as decreasing relative risk aversion (DRRA), increasing relative risk aversion (IRRA), and constant relative risk aversion (CRRA). In recent years, the HARA utility function has been applied in various economic models. Achury et al. [26] study a two-asset portfolio choice model with the HARA utility function which includes the subsistence consumption parameter. Levaggi and Menoncin [27] consider optimal dynamic tax evasion with the HARA utility function and a penalty function. Menoncin and Nembrini [28] find a closed form solution for the stochastic continuous time growth model with the HARA utility function.

Some researchers show that the problem of wealth distribution is not only related to the optimal conditions of economic growth but also closely related to the wealth inequality and the formulation of tax policies. To reduce the inequality of wealth distribution, it is very important to analyze the factors that affect wealth distribution, such as the preferences of agents, trading strategies, and trading time. Motivated by the works in [10, 17, 21, 28], we assume that agents estimate the next saving rate and profit of assets based on past trading information. We allow agents to adopt certain trading strategies to maximize their utility, such as using part of the assets for trading and the rest for savings. Under these assumptions, we investigate the influence of agent’s trading behavior and average consumption level on the wealth distribution.

The objective of this work is to employ the HARA utility function to describe the agents’ behavior and discuss the influence of average consumption level and trading time of agents on wealth distribution in the multiagent market. Compared with the works in [17, 21], we utilize the HARA utility function to characterize the agent’s behavior and divide the agent’s transaction into multiple periods. A wealth distribution model on isolated discrete time domains, which allows the wealth to exchange on unevenly spaced time points, is given. HARA utility function and the wealth distribution model are different from those in [17, 21], in which the CRRA utility function is used and the time domain of the wealth distribution model is a uniformly discrete set of points which allows the wealth to exchange at uniform time interval. HARA utility function contains several kinds of utility functions, e.g., DRRA utility function, IRRA utility function, and CRRA utility function. The time domain of the wealth distribution model in our work is a collection of points along the real number, which describes the distribution of wealth in nonuniform time interval. Consequently, we extend parts of results in [17, 21].

The study is organized as follows. We introduce the calculus and dynamic equation on isolated discrete time domains in Section 2. The Euler equation and transversality condition of a nonlinear stochastic dynamic model with a bivariate utility function are derived in Section 3. The main results about the wealth distribution model are given in Section 4. In Section 5, a numerical example is given to illustrate the advantages of the wealth distribution model on isolated discrete time domains.
2. Calculus and Dynamic Equation on Isolated Discrete Time Domains

In Sections 3 and 4, we will set up the nonlinear stochastic dynamic model and the wealth distribution model on the isolated discrete time domains. The isolated discrete time domains are different from the classic discrete time domain. The classical discrete time domain contains only uniform time points, while the isolated discrete time domains contain not only uniform time points but also nonuniform time points. Therefore, we will introduce some definitions, theorems, properties, and dynamic equations on isolated discrete time domains.

In this section, the basic concepts about delta-derivative, delta-integral, exponential function, and dynamic equation on isolated discrete time domains are introduced.

Let $\mathbb{T}$ be any nonempty closed subset of the real numbers $\mathbb{R}$. The operators $\sigma(t)$, $p(t)$, and $\mu(t)$ are defined by $\sigma(t) = \inf \{ s \in \mathbb{T} : s > t \}$, $p(t) = \sup \{ s \in \mathbb{T} : s < t \}$, and $\mu(t) = \sigma(t) - t$, respectively, where $\sigma(t)$ is the forward jump operator, $\mu(t)$ is the forward graininess operator, and $p(t)$ is the backward jump operator. From the time scale $\mathbb{T}$, the set $\mathbb{T}^k$ is obtained. For a point $t \in \mathbb{T}$ if $\sigma(t) > t$, we say that $t$ is right-scattered, and if $p(t) < t$, we say that $t$ is left-scattered. When $\mathbb{T}$ has a left-scattered maximum $m$, we have $\mathbb{T}^k = \mathbb{T} - \{m\}$. Otherwise, $\mathbb{T}^k = \mathbb{T}$.

It is worth noting that, in the case $\mathbb{T} = \mathbb{R}$, we obtain (see Merrell et al. [29])

$$
\begin{align*}
\sigma(t) &= p(t) = t, \\
\mu(t) &= 0. 
\end{align*}
$$

When $\mathbb{T} = \mathbb{Z}$, where $\mathbb{Z}$ is the set of positive integers, for any $t \in \mathbb{Z}$, we have

$$
\begin{align*}
\sigma(t) &= t + 1, \\
p(t) &= t - 1, \\
\mu(t) &= 1. 
\end{align*}
$$

When $\mathbb{P}^n = \{P^j \mid n \in \mathbb{N}\}$, where $P > 1$ and $\mathbb{N}$ is the natural number, we get

$$
\begin{align*}
\sigma(t) &= Pt, \\
p(t) &= \frac{t}{P}, \\
\mu(t) &= (P - 1)t. 
\end{align*}
$$

$\mathbb{P}^n = \{1, P, P^2, \ldots\}$ is the called quantum time scale, which is very important in quantum theory (see Merrell et al. [29]).

Let $F: \mathbb{T} \rightarrow \mathbb{R}$ be a real-valued function on $\mathbb{T}$. The delta-derivative [21, 30] of $F$ is denoted by

$$
F^\Delta(t) = \frac{F(\sigma(t)) - F(t)}{\mu(t)},
$$

where $t \in \mathbb{T}$. The delta-integral [21, 30] of $F$ is defined by

$$
\int_0^T F(t) \Delta t = \sum_{s \in (0,T) \cap \mathbb{T}} \mu(s) F(s),
$$

and the exponential function [21, 25, 30] is given as

$$
e_q(t, t_0) = \prod_{s \in \{t_0, t\} \cap \mathbb{T}} (1 + \mu(s)q(s)),
$$

where $1 + q(t)\mu(t) \neq 0$ for all $t \in \mathbb{T}$.

We mention that when $\mathbb{T} = \mathbb{Z}$, exponential function $e_q(t, t_0)$ becomes

$$
e_q(t, t_0) = (1 + q)^{t-t_0},
$$

where $1 + q \neq 0$.

When $\mathbb{T} = \mathbb{P}^n$, we get

$$
e_q(t, t_0) = \prod_{s \in \{t_0, t\} \cap \mathbb{T}} [1 + (P-1)qs], \quad t > t_0.
$$

For exponential function $e_{(q-1)/p}(t, t_0)$, we have

$$
e_{(q-1)/p}(t_0) = \prod_{s \in \{t_0, t\} \cap \mathbb{T}} \left( \frac{1}{q(s)} \right), \quad t < t_0.
$$

Let $0 < q < 1$ be a constant and $n_i$ be a function of $t$, where $n_i$ indicates that the number of discrete points on the interval $[t_0, t]$. Exponential function $e_{(q-1)/p}(t, t_0)$ is rewritten by

$$
e_{(q-1)/p}(t, t_0) = \prod_{s \in \{t_0, t\} \cap \mathbb{T}} q(s) = q^{n_0},
$$

where

$$
n_i(t, s) = \int_t^s \Delta(t) \frac{\mu(t)}{q(s)},
$$

To calculate the wealth distribution model in Section 4, we introduce an algorithm on time scales.

**Lemma 1** (see [30]). Let $Q$ be a $n \times n$ matrix. If $\lambda_1, \lambda_2, \lambda_3, \ldots, \lambda_n$ are the eigenvalues of $Q$, then

$$
Q^T = \sum_{i=0}^{n-1} r_{i+1}(t) U_i,
$$

where $r_i(t)$ ($i = 1, 2, \ldots, n$) are chosen to satisfy the system

$$
\begin{bmatrix}
r_1(t + 1) \\
r_2(t + 1) \\
\vdots \\
r_n(t + 1)
\end{bmatrix}
= 
\begin{bmatrix}
\lambda_1 & 0 & 0 & \cdots & 0 \\
1 & \lambda_2 & 0 & \cdots & 0 \\
0 & 1 & \lambda_3 & \cdots & 0 \\
\vdots & \vdots & \vdots & \ddots & \vdots \\
0 & \cdots & 0 & \cdots & \lambda_n
\end{bmatrix}
\begin{bmatrix}
r_1(t) \\
r_2(t) \\
\vdots \\
r_n(t)
\end{bmatrix}
$$

(13)
and $U_i$ are defined by
\[
U_0 = I, \\
U_i = (Q - \lambda_i I)U_{i-1}, \quad (1 \leq i \leq n).
\] (14)

The algorithm in Lemma 1 is called Putzer algorithm. The application of Lemma 1 is displayed in the numerical example.

**Definition 1** (see [30]). If $h \in \mathcal{R}$ and $f: \mathbb{T} \rightarrow \mathbb{R}$ is the right-dense continuous function, then the dynamic equation
\[
y(t) = h(t)y(t) + f(t),
\] (15)

is called regressive.

**Definition 2** (see [30]). Let $h: \mathbb{T} \rightarrow \mathbb{R}$ be the right-dense continuous function. Then, the circle minus of $h(t)$ for all $t \in \mathbb{T}^h$ is denoted by
\[
\Theta h(t) := \frac{h(t)}{1 + \mu(t)h(t)},
\] (16)

where $1 + \mu(t)h(t) \neq 0$, $h \in \mathcal{R}$ is the set of all regressive.

**Lemma 2** (see [21]). Let $Y_t$ be an $n \times 1$ vector function on isolated discrete time domains and satisfy
\[
Y_t = AE_t [Y_{\ast}^T] + F(t, Z_t),
\] (17)

where $A$ is the $n \times n$ nonsingular matrix and $F$ is the $n \times 1$ vector function. Then, the solution of equation (17) is
\[
Y_t = e_{(I-A)A^{-1}(1/\rho)}(t, 0)M(t) - e_{(I-A)A^{-1}(1/\rho)}(t, 0) \int_0^T e_{\Theta(I-A)A^{-1}(1/\rho)}(s, 0) \frac{1}{\mu(s)}F(s, Z_s)\Delta s,
\] (18)

where $t \in \mathbb{T}$ and $M(t)$ is a $n \times n$ arbitrary matrix martingale on $\mathbb{T}$ and meets the martingale property $E_t[M^\circ(t)] = M(t)$.

For more information of dynamic equations on isolated discrete time domains, see [30].

### 3. Nonlinear Stochastic Dynamic Model with a Bivariate Utility Function

We assume that the isolated discrete time domains $\mathbb{T}$ satisfy sup $\mathbb{T} = \infty$. Let $\mathbb{T} \cap [0, \infty) = [0, \infty)$. Then, the nonlinear stochastic dynamic model (NSDM) with a bivariate utility function on $\mathbb{T}$ is given by

\[
sup_{\{c_{i,t}, a_{\sigma(t)}\}_{t \in \mathbb{T}}} E_0 \sum_{t \in \{0, \infty\}} e^{(\beta_{i,t})\mu(t)}U(c_{i,t}, \tau_t)\mu(t) \tag{19}
\]

\[
\text{s.t.}
\begin{align*}
\tau_t &= f(a_{i,t}, a_{\sigma(t)}, y_{i,t}, y_{\sigma(t)}, r_{i,t}, r_{\sigma(t)}), \\
a_{i,t}, a_{\sigma(t)}, \tau_t, & \in X, \quad \text{are given},
\end{align*}
\]

where $0 < \beta_{i,t} < 1$, $a_{i,t}$ ($i = 1, 2, \ldots, n$), and $\sigma_{\sigma(t)}$ are the state variables, $\tau_t$ is the forward jump operator, $y_{i,t}$ ($i = 1, 2, \ldots, n$), and $f(x)$ is the random state variables, $E_0$ represents the mathematical expectation, $U: \mathbb{R}^2 \rightarrow \mathbb{R}^\ast$ is a strictly increasing, concave, continuous, and differentiable utility function, $X = \kappa \times \kappa$, where $\kappa = \{x_{1,\infty}: x_{1,\infty} = 0, t \in [0, \infty)\}$, and $f(x, y, z, u): \mathbb{R}^4 \rightarrow \mathbb{R}^\ast$ is a concave in $(x, y, z, u)$, continuous, and differentiable real-valued function with $f \neq 0$. The discount factor relating to NSDM (19) is $e^{(\beta_{i,0})\mu(t)}$. When $\mathbb{T} = \mathbb{Z}$, $e^{(\beta_{i,0})\mu(t)} = \beta_i^\ast$, where $\beta_i^\ast$ is a standard discount factor in the case of uniform time domain.

We mention that the concavity in $(x, y)$ means concavity in two variables $x$ and $y$ jointly, not in each variable separately. The concavity of $f(x, y, z, u)$ is used to find the optimal solution of NSDM (19), namely, function $f(x, y, z, u)$ is concave in $(x, y)$ if and only if for any pair of distinct points $(x, y, z, u)$ and $(\bar{x}, \bar{y}, \bar{z}, \bar{u})$ in the domain, we have

\[
f(x, y, z, u) - f(x^\ast, y^\ast, z, u) \leq f_x(x^\ast, y^\ast, z, u)(x - x^\ast)
\]

\[
+ f_y(x^\ast, y^\ast, z, u)(y - y^\ast),
\] (20)

\[
f(\bar{x}, \bar{y}, \bar{z}, \bar{u}) - f(x^\ast, \bar{y}^\ast, \bar{z}, \bar{u}) \leq f_x(x^\ast, \bar{y}^\ast, \bar{z}, \bar{u})(\bar{x} - x^\ast)
\]

\[
+ f_y(x^\ast, \bar{y}^\ast, \bar{z}, \bar{u})(\bar{y} - \bar{y}^\ast).
\] (21)

To find optimal sequence $\{c_{i,t}, a_{\sigma(t)}\}_{t = 0}^{\infty}$ to maximize the expected utility in NSDM (19), the Euler equation and transversality condition for NSDM (19) need to be given.

An Euler equation is the first-order condition for describing optimal choice of dynamic stochastic problems. There are several methods for the analysis and definition of Euler equation, namely, calculus of variations, optimal control theory, the Lagrangian method, and dynamic programming [31, 32]. We assume that NSDM (19) reaches the upper bound at $\{c_{i,t}, a_{\sigma(t)}\}_{t = 0}^{\infty}$. The Euler equation for NSDM (19) is defined as

\[
\mu U'_{i,t} f'(a_{i,t}, a_{\sigma(t)}, y_{i,t}, y_{\sigma(t)}, r_{i,t}, r_{\sigma(t)}) + E_t \left[ \beta_0 \mu_{\sigma(t)} U'_{i,t} f'(a_{i,t}, a_{\sigma(t+1)}, y_{i,t+1}, y_{\sigma(t+1)}, r_{i,t+1}, r_{\sigma(t+1)}) \right] = 0,
\] (22)
where $E_t$ represents the conditional expectation on the available information at time $t$, and the information is included in $I_t$, which contains the observations on $c_{i,t}, y_{i,t}, r_{i,t}$, and their past values; $f_x$ and $f_y$ are partial derivatives of function $f$ which is given in NSDM (19) with respect to the first and second variables, respectively.

The transversality condition and Euler equation are often used to describe the optimal paths of dynamic economic models. There are two main explanations for the transversality condition: (1) when the marginal utility is greater than 0, agents have the possibility of improving welfare near the end of life, and wealth is completely used for consumption. (2) When the marginal utility is equal to 0, the welfare of the agents cannot be improved. Usually, when the end point is fixed, there are several paths satisfying the Euler equation. The purpose of the transversality condition is to select the optimal path from the paths satisfying the Euler equation.

The transversality condition for NSDM (19) is given by

$$
\text{From Euler equation (22) and transversality condition (23), the optimal solution of NSDM (19) is found. In other words, Euler equation (22) and transversality condition (23) are sufficient for the existence of the optimal solution of NSDM (19).}
$$

**Theorem 1.** If sequence $\{c_{i,t}, a_{i,t}\}_{t=0}^{\infty}$ meets systems (22) and (23), then it is the optimal sequence for NSDM (19).

$$
\text{(NSDM)*} - \text{(NSDM)} = E_0 \lim_{T \to \infty} \left\{ \sum_{t \in [0, T]} e^{(\beta-1)T} \mu(t) \left[ U^*_{c_{i,t}}( \tau_i^* ) - U^*_{c_{i,t}}( \tau_i ) \right] \right\} \geq 0,
$$

where $(\text{NSDM)*}$ and $(\text{NSDM)}$ are the values of the objective function of NSDM (19) at $\{c_{i,t}, a_{i,t}\}_{t=0}^{\infty}$ and $\{c_{i,t}, a_{i,t}\}_{t=0}^{\infty}$, respectively.

Using the Dominated Convergence Theorem [33], we get

From equations (20) and (21), we obtain

$$
c_{i,t}^* - c_{i,t} \geq f_x(\alpha_{i,t}, a_{i,t}^*, y_{i,t}, r_{i,t})(a_{i,t}^* - a_{i,t}) + f_y(\alpha_{i,t}, a_{i,t}^*, y_{i,t}, r_{i,t})(a_{i,t}^* - a_{i,t}), \tag{26}
$$

$$
\tau_i^* - \tau_i \geq f_x(\alpha_{i,t}^*, \alpha_{i,t}, y_{i,t}, r_{i,t})(a_{i,t}^* - a_{i,t}) + f_y(\alpha_{i,t}^*, \alpha_{i,t}, y_{i,t}, r_{i,t})(a_{i,t}^* - a_{i,t}). \tag{27}
$$

Substituting equations (26) and (27) into inequality (25), we have
\[(NSDM)^* - (NSDM) \geq \lim_{T \to \infty} E_0 \sum_{t \in [0, \sigma(T)] T} e^{\rho_{(t, 0)}(t, 0) \mu_t U^T_{c_T}} \{ f_s (a^*_{i,t}, a^*_{i,\sigma(t)}, y_{i,\sigma(t)}, r_{i,\sigma(t)}) (a^*_{i,t} - a_{i,t}) + f_y (a^*_{i,t}, a^*_{i,\sigma(t)}, y_{i,\sigma(t)}, r_{i,\sigma(t)}) (a^*_{i,t} - a_{i,t}) \} \ + \lim_{T \to \infty} E_0 \sum_{t \in [0, \sigma(T)] T} e^{\rho_{(t, 0)}(t, 0) \mu_t U^T_{c_T}} \cdot \{ f_s (\pi^*_{i}, \pi^*_{\sigma(t)}, \pi^*_{\sigma(t)}, \pi_{\sigma(t)}) (\pi^*_{i} - \pi_{i}) + f_y (\pi^*_{i}, \pi^*_{\sigma(t)}, \pi^*_{\sigma(t)}, \pi_{\sigma(t)}) (\pi^*_{i} - \pi_{i}) \} \]

Utilizing the concavity assumption on the function \(f(x, y, z, u)\), we obtain

\[
I = \lim_{T \to \infty} E_0 \left[ e^{\rho_{(T, 0)}(T, 0) \mu_T U^T_{c_T}} \{ f_s (a^*_{i,T}, a^*_{i,\sigma(T)}, y_{i,\sigma(T)}, r_{i,\sigma(T)}) (a^*_{i,T} - a_{i,T}) + e^{\rho_{(T, 0)}(T, 0) \mu_T U^T_{c_T}} \cdot \{ f_s (\pi^*_{i,T}, \pi^*_{\sigma(T)}, \pi^*_{\sigma(T)}, \pi_{\sigma(T)}) (\pi^*_{i,T} - \pi_{i,T}) \} \right]
\]

Since \(a_{i,0} = a^*_{i,0}\) and \(\pi_0 = \pi^0\) at the initial state, equation (29) becomes

\[
I = \lim_{T \to \infty} \left[ e^{\rho_{(T, 0)}(T, 0) \mu_T U^T_{c_T}} \{ f_s (a^*_{i,T}, a^*_{i,\sigma(T)}, y_{i,\sigma(T)}, r_{i,\sigma(T)}) (a^*_{i,T} - a_{i,T}) + e^{\rho_{(T, 0)}(T, 0) \mu_T U^T_{c_T}} \cdot \{ f_s (\pi^*_{i,T}, \pi^*_{\sigma(T)}, \pi^*_{\sigma(T)}, \pi_{\sigma(T)}) (\pi^*_{i,T} - \pi_{i,T}) \} \right]
\]

From the recurrence relation of the exponential function, we obtain \(e^{\rho_{(t, 0)}(t, 0) \mu_t U^T_{c_T}}(t, 0) = \beta e^{\rho_{(t, 0)}(t, 0) \mu_t U^T_{c_T}}(t, 0)\). Consequently, equation (30) is turned into
$$I = \lim_{T \to \infty} E_0 \sum_{t \in \{0, T\}} \beta_t e^{(\beta_t - 1) \mu_t (t, T) \mu_T} \left\{ U_{\gamma(t)} f_x \left( a_{i, \sigma(t)}^*, a_{i, \sigma(t)}^* + y_{i, \sigma(t)}^*, r_{i, \sigma(t)}^* \right) \left( a_{i, \sigma(t)}^* - a_{i, \sigma(t)} \right) \\
+ U_{\gamma(t)} f_y \left( \bar{a}_{\sigma(t)}^*, \bar{a}_{\sigma(t)}^+ \bar{y}_{\sigma(t)}^*, \bar{r}_{\sigma(t)}^* \right) \left( \bar{a}_{\sigma(t)}^* - \bar{a}_{\sigma(t)} \right) \right\}$$

$$+ \lim_{T \to \infty} E_0 \sum_{t \in \{0, T\}} e^{(\beta_t - 1) \mu_t (t, T) \mu_T} \left\{ U_{\gamma(t)} f_x \left( a_{i, \sigma(t)}^*, a_{i, \sigma(t)}^* + y_{i, \sigma(t)}^*, r_{i, \sigma(t)}^* \right) \left( a_{i, \sigma(t)}^* - a_{i, \sigma(t)} \right) \\
+ U_{\gamma(t)} f_y \left( \bar{a}_{\sigma(t)}^*, \bar{a}_{\sigma(t)}^+ \bar{y}_{\sigma(t)}^*, \bar{r}_{\sigma(t)}^* \right) \left( \bar{a}_{\sigma(t)}^* - \bar{a}_{\sigma(t)} \right) \right\}$$

or

$$I = \lim_{T \to \infty} E_0 \sum_{t \in \{0, T\}} e^{(\beta_t - 1) \mu_t (t, T) \mu_T} \left\{ U_{\gamma(t)} f_x \left( a_{i, \sigma(t)}^*, a_{i, \sigma(t)}^* + y_{i, \sigma(t)}^*, r_{i, \sigma(t)}^* \right) \left( a_{i, \sigma(t)}^* - a_{i, \sigma(t)} \right) \\
+ U_{\gamma(t)} f_y \left( \bar{a}_{\sigma(t)}^*, \bar{a}_{\sigma(t)}^+ \bar{y}_{\sigma(t)}^*, \bar{r}_{\sigma(t)}^* \right) \left( \bar{a}_{\sigma(t)}^* - \bar{a}_{\sigma(t)} \right) \right\}$$

From system (22), we get

$$I = 0 + \lim_{T \to \infty} E_0 e^{(\beta_t - 1) \mu_t (t, T) \mu_T} \left\{ U_{\gamma(t)} f_x \left( a_{i, \sigma(t)}^*, a_{i, \sigma(t)}^* + y_{i, \sigma(t)}^*, r_{i, \sigma(t)}^* \right) \left( a_{i, \sigma(t)}^* - a_{i, \sigma(t)} \right) + U_{\gamma(t)} f_y \left( \bar{a}_{\sigma(t)}^*, \bar{a}_{\sigma(t)}^+ \bar{y}_{\sigma(t)}^*, \bar{r}_{\sigma(t)}^* \right) \left( \bar{a}_{\sigma(t)}^* - \bar{a}_{\sigma(t)} \right) \right\}.$$  

Employing system (23) and \( f_y \leq 0 \), we obtain

$$I \geq \lim_{T \to \infty} E_0 e^{(\beta_t - 1) \mu_t (t, T) \mu_T} \left\{ U_{\gamma(t)} f_x \left( a_{i, \sigma(t)}^*, a_{i, \sigma(t)}^* + y_{i, \sigma(t)}^*, r_{i, \sigma(t)}^* \right) \left( a_{i, \sigma(t)}^* - a_{i, \sigma(t)} \right) + U_{\gamma(t)} f_y \left( \bar{a}_{\sigma(t)}^*, \bar{a}_{\sigma(t)}^+ \bar{y}_{\sigma(t)}^*, \bar{r}_{\sigma(t)}^* \right) \left( \bar{a}_{\sigma(t)}^* - \bar{a}_{\sigma(t)} \right) \right\} = 0.$$  

Thus, \((\text{NSDM})^* - (\text{NSDM}) \geq 0\). This proof is completed.

4. The Wealth Distribution Model with the HARA Utility Function

Inspired by the works in [17, 21], the wealth distribution model with the HARA utility function is given by
where $E_0$ is the conditional expectation, $c_{it}$ $(i = 1, 2, \ldots, n)$ denotes the consumption of agent $i$ at time $t$, $\tau$, represents the average consumption of agents in the economy at time $t$, $0 < \varphi < 1$, $0 < \theta < 1$, $0 < \phi_i < 1$, $0 < \gamma_i < 1$ $(i = 1, 2, \ldots, n)$, all agents’ consumption in the economy satisfies $c_{it} > y_{it}$ by limiting the value of $y_{it}$, $0 < \lambda < 1$ is the saving rate of agent $i$, $a_{it}$ $(i = 1, 2, \ldots, n)$ denotes the wealth of agent $i$ at time $t$, $\tau$, is the average wealth of agents at time $t$, $y_{it}$ $(i = 1, 2, \ldots, n)$ represents the earnings of agent $i$ at time $t$, $\tau$, is the average income of agents at time $t$, $r_t$ denotes the rate of return of agents on wealth process at time $t$, all agents are assumed to have the same rate of return at the same time, and the discount factor $0 < \beta_i < 1$ $(i = 1, 2, \ldots, n)$ and the relative risk aversion parameter $0 < \delta_i < 1$ $(i = 1, 2, \ldots, n)$ are different to each agent $i$ and evolve stochastically.

In model (34), the wealth of agent $i$ in time $t$ is divided into two parts. Part of the wealth comes from the income that the agent can use part of the remaining wealth in time $t - 1$ for savings, that is, $(1 + r_t)\lambda(a_{it-1} - c_{it-1})$, where $a_{it-1} - c_{it-1}$ is the remaining wealth of the previous. The second part of the wealth comes from market transaction, namely, $y_{it} - (1 - \lambda)(a_{it-1} - c_{it-1})$, where $y_{it}$ is the wealth obtained by agent $i$ in the transaction, and $(1 - \lambda)(a_{it-1} - c_{it-1})$ indicates that agent $i$ uses part of the remaining wealth in time $t - 1$ for trading.

In model (34), agents are allowed to use part of the remaining wealth at time $t - 1$ for trading and part for savings. Agents adopt specific trading strategies that determine whether the agent gains or loses in wealth exchange to maximize their utility. In model (34), $y_{it} - (1 - \lambda)(a_{it-1} - c_{it-1}) > 0$ $(i = 1, 2, \ldots, n)$ denotes the net profits of agent $i$ from wealth exchange, and $y_{it} - (1 - \lambda)(a_{it-1} - c_{it-1}) < 0$ $(i = 1, 2, \ldots, n)$ represents the amount of wealth lost by agent $i$ in the process of wealth exchange. When $y_{it} - (1 - \lambda)(a_{it-1} - c_{it-1}) = 0$ $(i = 1, 2, \ldots, n)$ in model (34), agent $i$ does not make profits in the process of wealth exchange.

The utility function of model (34) is

$$U(c_{it}, \tau_t) = \frac{[c_{it} - y_{it}]^{1 - \delta_t} - 1}{1 - \delta_t},$$

which describes the preference of agent $i$ for risk. Equation (35) is so-called the HARA utility function which captures the idea of “keeping up with the Joneses,” namely, the consumption standards of agents are affected by average consumption level. When time $t$ is fixed, $\tau_t$ becomes a constant parameter $\tau$ which plays an important role in determining agent’s preferences.

(i) When $\tau > 0$, utility function (35) describes an agent with decreasing relative risk aversion (DRRA) preference, and $\tau$ is interpreted as the minimum standard of living for consumption. In utility function (35), when $c_{it} \to \tau$, the marginal utility tends to infinity, and the optimal consumption level will never reach $\tau = \tau$ and always satisfies $c_{it} > \tau$ (see Achury et al. [26]).

(ii) When $\tau < 0$, utility function (35) characterizes an agent with increasing relative risk aversion (IRRA) preference. $\tau$ in utility function (35) indicates that agents become more conservative as their consumption increases. However, $\tau$ is no longer interpreted as the agent’s minimum consumption level. Levaggi and Menoncin [27] introduce the increasing relative risk aversion preference.

(iii) When $\tau = 0$, utility function (35) describes an agent with constant relative risk aversion preference which is the CRRA utility function.

We set the isolated discrete time domains $T$ to satisfy $\sup T = \infty$. Let $T \cap [0, \infty) = [0, \infty)$. From model (34), the wealth distribution model on isolated discrete time domains $T$ has the following form:
Remark 1. When $\mathbb{T} = \mathbb{Z}$, wealth distribution model (36) becomes model (34). Consequently, model (36) is a generalization of model (34).

$$\sup_{\left\{ c_{i,t}, a_{i,t} \right\}_{t=0}^{\infty}} E_0 \sum_{t=0}^{\infty} \left[ e^{(\beta_{-1})_q(t)(T,0)} \left( \frac{c_{i,t} - y_i \tau_{i,t}}{1 - \delta_i} - \mu_t \right) \right],$$

where $\delta_i \neq 1, 2\lambda + \lambda r_{\tau}(t) - 1 \neq 0 \ (i = 1, 2, \ldots, n)$.

Consequently, the first-order condition of model (36) is

$$\begin{align*}
E_t \left[ \mu_t \left( c_{i,t} - y_i \tau_{i,t} \right)^{-\delta} \left( -\frac{1}{2\lambda + \lambda r_{\tau}(t) - 1} \right) + \beta_{\tau}(t) \left( c_{i,t} - y_i \tau_{i,t} \right)^{-\delta} \right] &= 0, \\
E_t \left[ \mu_t \left( c_{i,t} - \bar{y} \tau_{i,t} \right)^{-\delta} \left( -\frac{\bar{y}}{2\lambda + \lambda r_{\tau}(t) - 1} \right) - \beta_{\bar{y}}(t) \left( c_{i,t} - \bar{y} \tau_{i,t} \right)^{-\delta} \right] &= 0,
\end{align*}$$

Employing system (38), the transversality condition of model (36) becomes

$$\begin{align*}
\lim_{T \to \infty} E_0 e^{(\beta_{-1})_q(t)(T,0)\mu_t \left( c_{i,t} - y_i \tau_{i,t} \right)^{-\delta} \left( -\frac{1}{k_{\tau}(T)} \right) a_{i,\tau(t)}} &= 0, \\
\lim_{T \to \infty} E_0 e^{(\beta_{-1})_q(t)(T,0)\mu_t \left( c_{i,t} - \bar{y} \tau_{i,t} \right)^{-\delta} \left( \frac{\bar{y}}{k_{\bar{y}}(T)} \right) \bar{y} \tau_{i,t}}} &= 0.
\end{align*}$$

Step II. Expression of the initial state.
To seek out the initial state, the constraints of model (36) are rewritten as follows:
\[ c_{i,t} = a_{i,t} - \frac{1}{2\lambda + \lambda r_{\sigma(t)}} - \frac{1}{2\lambda + \lambda r_{\sigma(t)}} - Y_{i,\sigma(t)}, \]  
(41a)\\
\[ \bar{c}_t = \bar{\alpha}_t - \frac{1}{2\lambda + \lambda r_{\sigma(t)}} - \frac{1}{2\lambda + \lambda r_{\sigma(t)}} - Y_{i,\sigma(t)}, \]  
(41b)\\
\[ k_t = 2\lambda + \lambda r_\lambda - 1, \]  
(41c)\\
\[ \log y_{i,t} = (1 - \phi_t) \log y_{i,0} + \phi_t \log y_{i,t} + \epsilon_{i,t}, \]  
(41d)\\
\[ \log \bar{y}_t = (1 - \bar{\phi}) \log \bar{y}_0 + \bar{\phi} \log \bar{y}_t + \bar{\epsilon}_t, \]  
(41e)\\
\[ \log r_t = \varphi \log r_0 + (1 - \varphi) \log r_{t} + \eta_t. \]  
(41f)

From system (39) and equations (41a)–(41f), the initial states of model (36) are
\[ c_{i,0} = a_{i,0} - \frac{1}{2\lambda + \lambda r_{\sigma(0)}} - \frac{1}{2\lambda + \lambda r_{\sigma(0)}} - Y_{i,\sigma(0)}, \]  
(42a)\\
\[ \bar{c}_0 = \bar{\alpha}_0 - \frac{1}{2\lambda + \lambda r_{\sigma(0)}} - \frac{1}{2\lambda + \lambda r_{\sigma(0)}} - Y_{i,\sigma(0)}, \]  
(42b)\\
\[ k_0 = 2\lambda + \lambda r_0 - 1, \]  
(42c)\\
\[ E_0 \left[ \beta \frac{\mu_{\sigma(0)}}{\mu_0} \left( \frac{c_{i,0} - Y_{i} \bar{c}_0}{c_{i,\sigma(0)} - Y_{i} \bar{c}_0} \right) k_{\sigma(0)} \right] = 1, \]  
(42d)\\
\[ E_0 \left[ \beta \frac{\mu_{\sigma(0)}}{\mu_0} \left( \frac{c_{i,0} - Y_{i} \bar{c}_0}{c_{i,\sigma(0)} - Y_{i} \bar{c}_0} \right) k_{\sigma(0)} \right] = 1, \]  
(42e)\\
\[ \log y_{i,0} = (1 - \phi_0) \log y_{i,0} + \phi_0 \log y_{i,0} + \epsilon_{i,0}, \]  
(42f)\\
\[ \log \bar{y}_0 = (1 - \phi) \log \bar{y}_0 + \phi \log \bar{y}_0 + \bar{\epsilon}_0, \]  
(42g)\\
\[ \log r_0 = \varphi \log r_0 + (1 - \varphi) \log r_0 + \eta_0, \]  
(42h)

where \(2\lambda + \lambda r_{\sigma(0)} - 1 \neq 0\) \((i = 1, 2, \ldots, n), c_{i,\sigma(0)} - Y_{i} \bar{c}_0 \neq 0\) \((i = 1, 2, \ldots, n), c_{i,\sigma(0)} - Y_{i} \bar{c}_0 \neq 0, \) and \(\mu_0 \neq 0.\)

**Step III.** Log-linearization of the constraints and the Euler equation.

Let \(\bar{x}_t = (x_t - x_0)/x_0.\) To log-linearize equation (41a), our idea is to linearize each part of the equation and then sum it up. For \(c_{i,t}\), we have
\[ c_{i,t} \approx c_{i,0} (1 + \bar{c}_{i,t}). \]  
(43)

Through a series of algebraic calculations, we get (see Appendix A)
\[ \bar{c}_{i,t} = \frac{a_{i,0} - a_{i,\sigma(0)} - a_{i,\sigma(0)} - Y_{i,\sigma(0)}}{c_{i,0} k_{\sigma(0)} - Y_{i,\sigma(0)}}, \]  
(44)

where \(c_{i,0} \neq 0\) \((i = 1, 2, \ldots, n)\) and \(k_{\sigma(0)} \neq 0.\)

Similarly, we log-linearize equation (41b) and obtain
\[ \bar{c}_t \approx \frac{a_{0,0} - a_{0,\sigma(0)} - a_{0,\sigma(0)} - Y_{0,\sigma(0)}}{c_{0,0} k_{\sigma(0)} - Y_{0,\sigma(0)}}, \]  
(45)

where \(c_{i,0} \neq 0\) and \(k_{\sigma(0)} \neq 0.\)

We log-linearize two sides of equation (41c) and have
\[ k_t = k_0 (1 + \delta k_t), \]  
(46)

Dividing two sides of equation (46) by \(k_0\), we obtain
\[ \delta k_t = \frac{\lambda r_0 - \lambda r_0}{k_0}, \]  
(47)

where \(k_0 \neq 0.\)

Log-linearizing system (39), we get (see Appendix B)
\[ E_t \left[ \mu_{\sigma(t)} - \mu_t + \frac{c_{i,0} \delta c_t}{c_{i,\sigma(t)} - Y_{i} \bar{c}_t} c_{i,t} - \frac{c_{i,0} \delta c_t}{c_{i,\sigma(t)} - Y_{i} \bar{c}_t} c_{i,0} - Y_{i} \bar{c}_t \right] = 0, \]  
(48)\\
\[ E_t \left[ \mu_{\sigma(t)} - \mu_t + \frac{c_{i,0} \delta c_t}{c_{i,\sigma(t)} - Y_{i} \bar{c}_t} c_{i,t} - \frac{c_{i,0} \delta c_t}{c_{i,\sigma(t)} - Y_{i} \bar{c}_t} c_{i,0} - Y_{i} \bar{c}_t \right] = 0. \]  
(49)

Log-linearizing equation (41d), we obtain the following expression (see Appendix C):
\[ \bar{y}_{i,\sigma(t)} = \bar{\phi} \bar{y}_{i,t} + \bar{\phi} \bar{y}_{i,t} + \bar{\epsilon}_{i,\sigma(t)}, \]  
(50)

where \(\delta c_{i,0} \neq 0.\)

Similarly, we log-linearize equation (41e) and get
\[ \bar{y}_{\sigma(t)} = \bar{\phi} \bar{y}_{t} + \bar{\phi} \bar{y}_{t} + \bar{\epsilon}_{\sigma(t)}, \]  
(51)

where \(\bar{\epsilon}_{\sigma(t)} = 0.\)

Taking conditional expectation on both sides of equations (50) and (51) on information set \(I_t\) yields
\[ E_t \left[ \hat{y}_{i,\sigma(t)} \right] = \phi_t \bar{y}_t, \]  
\[ E_t \left[ \hat{\sigma}_{\varphi(t)} \right] = \bar{\sigma}_{\varphi_t}. \]  
(52)

Log-linearizing equation (41f), we obtain (see Appendix D)
\[ \bar{\varphi}_{\sigma(t)} \approx (1 - \varphi) \bar{\varphi}_t + \eta_{\sigma(t)} - \eta_{\sigma(0)}, \]  
(54)

where \( \eta_{\sigma(0)} = 0 \).

Adding conditional expectation on both sides of equation (54) on information set \( I_t \), we have
\[ E_t \left[ \bar{\varphi}_{\sigma(t)} \right] \approx (1 - \varphi) \bar{\varphi}_t. \]  
(55)

Step IV. Expressing the corresponding recursive equations by matrix operators.

Taking conditional expectation on both sides of equation (44) on information set \( I_t \), we get
\[ \frac{y_{i,\sigma(t)}}{c_{i,0} k_{\sigma(t)}} E_t \left[ \tilde{a}_{i,\sigma(t)} \right] = \frac{a_{i,0}}{c_{i,0}} a_{i,0} - \tilde{c}_{i,0} + \frac{y_{r,\sigma(t)}}{c_{i,0} k_{\sigma(t)}} E_t \left[ \bar{y}_{r,\sigma(t)} \right] 
+ \lambda_{i,\sigma(t)} r_{\sigma(t)} \sigma_{c(t)}^t \lambda_{i,0} - y_{i,\sigma(t)} E_t \left[ \sigma_{\varphi(t)} \right]. \]  
(56)

From equations (52) and (55), equation (56) becomes
\[ \frac{a_{i,\sigma(t)}}{c_{i,0} k_{\sigma(t)}} E_t \left[ \tilde{a}_{i,\sigma(t)} \right] = \frac{a_{i,0}}{c_{i,0}} a_{i,0} - \tilde{c}_{i,0} + \lambda_{i,\sigma(t)} r_{\sigma(t)} \sigma_{c(t)}^t \lambda_{i,0} - y_{i,\sigma(t)} \sigma_{\varphi(t)} \]  
\[ + \frac{a_{i,\sigma(t)}}{c_{i,0} k_{\sigma(t)}} \left( a_{i,0} - y_{i,\sigma(t)} \right) \left( 1 - \varphi \right) \bar{\varphi}_t. \]  
(57)

Likewise, equation (45) is rewritten by
\[ \frac{\eta_{\sigma(t)}}{c_{i,0} k_{\sigma(t)}} E_t \left[ \tilde{\varphi}_{\sigma(t)} \right] = \frac{\eta_{i,0}}{c_{i,0}} a_{i,0} - \tilde{c}_{i,0} + \lambda_{i,\sigma(t)} r_{\sigma(t)} \sigma_{c(t)}^t \lambda_{i,0} - y_{i,\sigma(t)} \sigma_{\varphi(t)} \]  
\[ + \frac{\eta_{\sigma(t)}}{c_{i,0} k_{\sigma(t)}} \left( a_{i,0} - y_{i,\sigma(t)} \right) \left( 1 - \varphi \right) \bar{\varphi}_t. \]  
(58)

From equations (47), (48), and (55), we have
\[ \tau_{i,\sigma(t)} \tau_{\sigma(t)} = \mu_{\sigma(t)} - \mu_t + \frac{c_{i,0} \delta_{i,0}}{c_{i,0} - \gamma tr_{t}^0} \tilde{c}_{i,0} \]  
\[ \frac{c_{i,0} \delta_{i,0}}{c_{i,0} - \gamma tr_{t}^0} \tilde{c}_{i,0} = \mu_{\sigma(t)} - \mu_t + \frac{c_{i,0} \delta_{i,0}}{c_{i,0} - \gamma tr_{t}^0} \tilde{c}_{i,0} \]  
(59)

Calculating equation (49) gives rise to
\[ \frac{c_{i,0} \delta_{i,0}}{c_{i,0} - \gamma tr_{t}^0} E_t \left[ \tilde{c}_{i,\sigma(t)} \right] = \frac{d tr_{t}^0}{c_{i,0} - \gamma tr_{t}^0} \tilde{c}_{i,0} \]  
\[ \frac{d tr_{t}^0}{c_{i,0} - \gamma tr_{t}^0} \tilde{c}_{i,0} = \mu_{\sigma(t)} - \mu_t + \frac{c_{i,0} \delta_{i,0}}{c_{i,0} - \gamma tr_{t}^0} \tilde{c}_{i,0} \]  
(60)

Equations (52), (53), (55), and (57)–(60) are rewritten by the matrix form, namely,
\[
\begin{bmatrix}
\tilde{a}_{i,\sigma(t)} \\
\tilde{a}_{\sigma(t)} \\
\tilde{c}_{i,\sigma(t)} \\
\tilde{c}_{\sigma(t)} \\
\tilde{y}_{i,\sigma(t)} \\
\tilde{y}_{\sigma(t)} \\
\tilde{\varphi}_{\sigma(t)} \\
\tilde{\varphi}_t
\end{bmatrix}
=
\begin{bmatrix}
\tilde{a}_{i,\sigma(t)} \\
\tilde{a}_t \\
\tilde{c}_{i,\sigma(t)} \\
\tilde{c}_t \\
\tilde{y}_{i,\sigma(t)} \\
\tilde{y}_t \\
\tilde{\varphi}_{\sigma(t)} \\
\tilde{\varphi}_t
\end{bmatrix}
+
\begin{bmatrix}
0 \\
\mu_{\sigma(t)} - \mu_t \\
\mu_{\sigma(t)} - \mu_t \\
0 \\
0 \\
0 \\
0 \\
0
\end{bmatrix}
\]  
(61)
where

\[
B = \begin{bmatrix}
\frac{a_{i,\sigma(0)}}{c_{i,0}k_{\sigma(0)}} & 0 & 0 & 0 & 0 & 0 \\
0 & \frac{\pi_{\sigma(0)}}{c_{i,0}k_{\sigma(0)}} & 0 & 0 & 0 & 0 \\
0 & 0 & \frac{c_{i,\sigma(0)}\delta_1}{c_{i,\sigma(0)} - \gamma_1c_{\sigma(0)}} & -\frac{c_{i,\sigma(0)}\delta_1}{c_{i,\sigma(0)} - \gamma_1c_{\sigma(0)}} & 0 & 0 \\
0 & 0 & \frac{c_{i,\sigma(0)}\delta}{c_{i,\sigma(0)} - \gamma c_{\sigma(0)}} & -\frac{c_{i,\sigma(0)}\delta}{c_{i,\sigma(0)} - \gamma c_{\sigma(0)}} & 0 & 0 \\
0 & 0 & 0 & 0 & 1 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 \\
\end{bmatrix}
\]  

\[
C = \begin{bmatrix}
\frac{a_{i,0}}{c_{i,0}} & 0 & -1 & 0 & \frac{y_{i,\sigma(0)}}{c_{i,0}k_{\sigma(0)}} & 0 & \frac{\lambda r_{\sigma(0)}}{k_{\sigma(0)}^{2}}W_{1}W \\
0 & \frac{\pi_{0}}{c_{0}} & 0 & -1 & 0 & \frac{\pi_{\sigma(0)}}{c_{\sigma(0)}k_{\sigma(0)}} & \frac{\lambda r_{\sigma(0)}}{k_{\sigma(0)}^{2}}W_{2}W \\
0 & 0 & \frac{c_{i,0}\delta_1}{c_{i,0} - \gamma_1c_{0}} & -\frac{\delta_1c_{0}}{c_{i,0} - \gamma_1c_{0}} & 0 & 0 & \frac{\lambda r_{\sigma(0)}}{k_{\sigma(0)}}W \\
0 & 0 & \frac{c_{i,0}\delta}{c_{i,0} - \gamma c_{0}} & -\frac{\delta c_{0}}{c_{i,0} - \gamma c_{0}} & 0 & 0 & \frac{\lambda r_{\sigma(0)}}{k_{\sigma(0)}}W \\
0 & 0 & 0 & 0 & \phi_{1} & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & \phi_{0} & 0 \\
\end{bmatrix}
\]

and \(c_{i,0} \neq 0, k_{\sigma(0)} \neq 0, c_{i,\sigma(0)} - \gamma_1c_{\sigma(0)} \neq 0, c_{i,\sigma(0)} - \gamma c_{\sigma(0)} \neq 0, \)
\(c_{i,0} - \gamma_1c_{0} \neq 0, c_{i,0} - \gamma c_{0} \neq 0 \) \((i = 1, 2, \ldots, n), c_{0}, \gamma, W = 1 - \varphi, W_{1} = a_{i,\sigma(0)} - \gamma_{i,\sigma(0)}, W_{2} = \pi_{\sigma(0)} - \gamma_{\sigma(0)}\) and \(k_{\sigma(0)}\) are nonzero constants.

**Step V.** Finding the optimal solution of equation (61).

**Theorem 2.** Let \(\mathbb{T}\) be the isolated discrete time domains. If matrix \(C\) is nonsingular, then the optimal solution of equation (61) is

\[
J_{t} = e_{(I-D)^{-1}(1/\mu)}(t, 0)M(t) - e_{(I-D)^{-1}(1/\mu)}(t, 0)\int_{0}^{T} e_{\Theta_{(I-D)^{-1}(1/\mu)}}(s, 0) \frac{1}{\mu(s)}F(s)\Delta s,
\]

where
Here, we state that the proof of Theorem 2 is similar to that of Lemma 2.3 in [21]. We omit its proof.

**Remark 3.** If $\mathbb{T} = \mathbb{Z}$, then solution (63) becomes

$$J_t = D^{-1} M(t) - D^{-1} \sum_{s=0}^{t-1} D^s F_s.$$  \hfill (65)

We mention that equation (65) is the optimal solution of model (34).

**Remark 4.** When $c_t = c$, $c_t$ indicates the average consumption level of agents during the whole trading period, and equation (61) becomes

$$B_1 E_t = C_1 \begin{bmatrix} \hat{a}_{i,t} \\ \hat{c}_{i,t} \\ \hat{y}_{i,t} \\ \hat{r}_{t} \end{bmatrix} + \begin{bmatrix} 0 \\ \mu_{t} \\ 0 \\ 0 \end{bmatrix},$$  \hfill (66)

where

and $M(t)$ is a $7 \times 1$ arbitrary matrix martingale on $\mathbb{T}$ and satisfies the martingale property $E_t[M^2(t)] = M(t)$.

We mention that equation (65) is the optimal solution of model (34).

**Remark 4.** When $\tau_t = \tau$, $\tau_t$ indicates the average consumption level of agents during the whole trading period, and equation (61) becomes

$$B_1 E_t = C_1 \begin{bmatrix} \hat{a}_{i,t} \\ \hat{c}_{i,t} \\ \hat{y}_{i,t} \\ \hat{r}_{t} \end{bmatrix} + \begin{bmatrix} 0 \\ \mu_{t} \\ 0 \\ 0 \end{bmatrix},$$  \hfill (66)

where

and $c_{i,t}(0) - y_i \tau \neq 0$. The optimal solution of equation (66) is

$$T_t = e^{(I-D_1)D_t^{-1}(1/\mu)(t,0)\tilde{M}(t)} - e^{(I-D_1)D_t^{-1}(1/\mu)(t,0)\int_0^T e^{\phi (I-D_1)D_s^{-1}(1/\mu)(s,0)\frac{1}{\mu(s)}F_1(s)\Delta s},$$  \hfill (68)
calculate the matrix and

\[
\tilde{M}(t) = \begin{bmatrix}
\tilde{a}_{i,t} \\
\tilde{c}_{i,t} \\
\tilde{y}_{i,t} \\
\tilde{F}_{i,t}
\end{bmatrix},
\]

and \( \tilde{M}(t) \) is a \( 4 \times 1 \) arbitrary matrix martingale and meets the martingale property \( E_t [\tilde{M}^r(t)] = \tilde{M}(t) \).

Remark 5. When \( \tau = 0 \), utility function (35) becomes the CRRA utility function which describes the agents with constant relative risk aversion preferences. Equation (66) becomes a wealth distribution model with the CRRA utility function.

From solution (63) of model (36), when the initial values \( a_{i,\sigma(0)}, c_{i,\sigma(0)}, y_{i,\sigma(0)}, \lambda_{i,\sigma(0)}, r_{\sigma(0)}, \tau_{\sigma(0)}, \pi_{\sigma(0)}, \tau_{\sigma(0)}, \) and \( k_{\sigma(0)} \) are fixed, the wealth distribution is only related to the parameters \( \sigma_\iota, \gamma_i, \sigma, \phi_i, \varphi, \) and \( \tilde{\lambda} \), namely, model (36) depicts the influence of agent’s trading behavior (the behavior of agent’s risk aversion and the behavior of forecasting market risk and saving tendency) on wealth distribution.

5. A Numerical Example

To explain solution (63), we choose parameter values from [21, 28] to calculate it. We set \( \gamma_i = 0.9, \lambda = 0.5, \delta = 0.2, \delta = 0.5, c_{i,0} = c_{i,\sigma(0)}, \tau_0 = \tau_{\sigma(0)} = 16, a_{i,0} = a_{i,\sigma(0)} = 50, \phi_0 = \phi_{\sigma(0)} = 0.98, \lambda_{i,\sigma(0)} = 0.32, \lambda_{i,\sigma(0)} = 0.9, r_{\sigma(0)} = 0.2, \lambda = 0.6, \phi = 0.05, \phi_1 = 0.85, \tilde{\phi} = 0.7, y_{i,\sigma(0)} = 20, \) and \( \tau_{\sigma(0)} = 15 \). Then, the matrices \( B, C, \) and \( D \) are obtained:

\[
B = \begin{bmatrix}
1.27 & 0 & 0 & 0 & 0 \\
0 & 5.86 & 0 & 0 & 0 \\
0 & 0 & 0.31 & -0.11 & 0 \\
0 & 0 & 0.625 & -0.125 & 0 \\
0 & 0 & 0 & 0 & 1 \\
0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0
\end{bmatrix},
\]

\[
C = \begin{bmatrix}
1.25 & 0 & -1 & 0 & 0.51 & 0 & 0.133 \\
0 & 1.875 & 0 & -1 & 0 & 1.17 & 1.04 \\
0 & 0 & 0.3125 & -0.1125 & 0 & 0 & 0.174 \\
0 & 0 & 0.625 & -0.125 & 0 & 0 & 0.116 \\
0 & 0 & 0 & 0 & 0.85 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0.07 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 0.95
\end{bmatrix},
\]

\[
D = \begin{bmatrix}
1.0160 & 0 & 0.8080 & -0.0080 & -0.4800 & 0 & 0.1224 \\
0 & 3.1253 & 0.0267 & 0.5067 & 0 & -8.9143 & 0.7186 \\
0 & 0 & 1.0100 & -0.0100 & 0 & 0 & 0.2931 \\
0 & 0 & 0.0500 & 0.9500 & 0 & 0 & 2.4421 \\
0 & 0 & 0 & 0 & 1.1765 & 0 & 0 \\
0 & 0 & 0 & 0 & 14.2857 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 1.0526
\end{bmatrix}.
\]

The Putzer algorithm on time scales [34] is employed to calculate the matrix \( e_{(1-D)D^{-1}(1/\mu)}(t,0) \). Consequently, solution (63) is rewritten as
\[ \tilde{a}_{12} = d_{11}M_1(t) + d_{12}M_2(t) + d_{13}M_3(t) + d_{14}M_4(t) + d_{15}M_5(t) + d_{16}M_6(t) \\
+ d_{17}M_7(t) - d_{13} \sum_{r \in \{0,7\}} g_{32}(\mu_{\sigma(t)} - \mu_t) - d_{14} \sum_{r \in \{0,7\}} g_{42}(\mu_{\sigma(t)} - \mu_t) \\
\tilde{a}_r = d_{31}M_1(t) + d_{21}M_2(t) + d_{32}M_3(t) + d_{23}M_4(t) + d_{33}M_5(t) + d_{24}M_6(t) \\
+ d_{27}M_7(t) - d_{33} \sum_{r \in \{0,7\}} g_{32}(\mu_{\sigma(t)} - \mu_t) - d_{34} \sum_{r \in \{0,7\}} g_{42}(\mu_{\sigma(t)} - \mu_t) \\
\tilde{c}_{12} = d_{31}M_1(t) + d_{32}M_2(t) + d_{33}M_3(t) + d_{34}M_4(t) + d_{35}M_5(t) + d_{36}M_6(t) \\
+ d_{37}M_7(t) - d_{33} \sum_{r \in \{0,7\}} g_{32}(\mu_{\sigma(t)} - \mu_t) - d_{34} \sum_{r \in \{0,7\}} g_{42}(\mu_{\sigma(t)} - \mu_t) \\
\tilde{c}_r = d_{41}M_1(t) + d_{42}M_2(t) + d_{43}M_3(t) + d_{44}M_4(t) + d_{45}M_5(t) + d_{46}M_6(t) \\
+ d_{47}M_7(t) - d_{13} \sum_{r \in \{0,7\}} g_{32}(\mu_{\sigma(t)} - \mu_t) - d_{44} \sum_{r \in \{0,7\}} g_{42}(\mu_{\sigma(t)} - \mu_t) \\
\tilde{y}_{12} = d_{51}M_1(t) + d_{52}M_2(t) + d_{53}M_3(t) + d_{54}M_4(t) + d_{55}M_5(t) + d_{56}M_6(t) \\
+ d_{57}M_7(t) - d_{53} \sum_{r \in \{0,7\}} g_{32}(\mu_{\sigma(t)} - \mu_t) - d_{54} \sum_{r \in \{0,7\}} g_{42}(\mu_{\sigma(t)} - \mu_t) \\
\tilde{y}_r = d_{61}M_1(t) + d_{62}M_2(t) + d_{63}M_3(t) + d_{64}M_4(t) + d_{65}M_5(t) + d_{66}M_6(t) \\
+ d_{67}M_7(t) - d_{63} \sum_{r \in \{0,7\}} g_{32}(\mu_{\sigma(t)} - \mu_t) - d_{64} \sum_{r \in \{0,7\}} g_{42}(\mu_{\sigma(t)} - \mu_t) \\
\tilde{r}_r = d_{71}M_1(t) + d_{72}M_2(t) + d_{73}M_3(t) + d_{74}M_4(t) + d_{75}M_5(t) + d_{76}M_6(t) \\
+ d_{77}M_7(t) - d_{73} \sum_{r \in \{0,7\}} g_{32}(\mu_{\sigma(t)} - \mu_t) - d_{74} \sum_{r \in \{0,7\}} g_{42}(\mu_{\sigma(t)} - \mu_t) \\
\] (70)

where \( M_1(t), M_2(t), M_3(t), M_4(t), M_5(t), M_6(t), \) and \( M_7(t) \) are arbitrary martingales and

\[
d_{11} = \lambda_1^n, \\
d_{12} = 0, \\
d_{16} = 0, \\
d_{13} = -66.49\lambda_1^n - 3.69 \times 10^{-3} \lambda_2^n - 3.83\lambda_4^n - 2.77 \times 10^{-15} \lambda_5^n \\
- 8.39 \times 10^{-24} \lambda_6^n + 6.74 \times 10^{-18} \lambda_7^n + 70.33, \\
d_{14} = 10.23\lambda_1^n + 3.64 \times 10^{-3} \lambda_2^n + 3.84\lambda_4^n + 9.57 \times 10^{-17} \lambda_5^n - 2.50 \times 10^{-23} \lambda_6^n - 2.081 \times 10^{-18} \lambda_7^n - 14.07, \\
d_{15} = -10.82\lambda_1^n + 0.05\lambda_2^n + 0.001\lambda_4^n + 4.06\lambda_5^n + 1.37 \times 10^{-23} \lambda_6^n + 1.7 \times 10^{-17} \lambda_7^n + 6.701, \\
d_{17} = -253.22\lambda_1^n - 0.12\lambda_2^n - 220.75\lambda_4^n - 13.13\lambda_5^n + 3.301 \times 10^{-7} \lambda_6^n + 0.031\lambda_7^n + 1365.98, \\
d_{21} = 0, \\
d_{22} = \lambda_2^n - 1.913 \times 10^{-5} \lambda_1^n, \\
d_{25} = 0, \\
d_{23} = -0.04\lambda_1^n + 0.017\lambda_2^n - 0.33\lambda_4^n - 1.22 \times 10^{-13} \lambda_5^n + 2.11 \times 10^{-19} \lambda_6^n \\
+ 3.15 \times 10^{-16} \lambda_7^n + 0.35, \\
d_{24} = -0.49\lambda_1^n + 0.23\lambda_2^n + 0.33\lambda_4^n - 2.87 \times 10^{-14} \lambda_5^n + 4.95 \times 10^{-20} \lambda_6^n \\
+ 7.43 \times 10^{-17} \lambda_7^n - 0.07, \\
\]
\[ \begin{align*}
d_{26} &= 13081293.87\lambda_1^n + 26451.40\lambda_2^n + 4751069.56\lambda_4^n + 336.031\lambda_5^n \\
&\quad - 28057.56\lambda_6^n - 9.65 \times 10^{-10}\lambda_7^n - 17831521.75, \\
d_{27} &= 27.45\lambda_1^n + 0.63\lambda_2^n + 27.55\lambda_4^n + 3.51\lambda_5^n - 8.83 \times 10^{-8}\lambda_6^n \\
&\quad - 0.008\lambda_7^n - 294.10, \\
d_{31} &= 0, \\
d_{32} &= 0, \\
d_{35} &= 0, \\
d_{36} &= 0, \\
d_{33} &= 2.13\lambda_1^n + 0.000265\lambda_2^n + 0.28\lambda_4^n + 2.44 \times 10^{-17}\lambda_5^n - 1.58 \times 10^{-24}\lambda_6^n \\
&\quad - 578\lambda_7^n - 1.41, \\
d_{34} &= -0.0013\lambda_1^n - 0.00026\lambda_2^n - 0.28\lambda_4^n + 5.02 \times 10^{-17}\lambda_5^n + 2.21 \times 10^{-24}\lambda_6^n \\
&\quad - 1.24 \times 10^{-19}\lambda_7^n + 0.28, \\
d_{37} &= -5.26\lambda_1^n + 0.00444\lambda_2^n + 0.022\lambda_4^n - 0.64\lambda_5^n + 1.62 \times 10^{-8}\lambda_6^n \\
&\quad + 0.0015\lambda_7^n + 49.02, \\
d_{41} &= 0, \\
d_{42} &= 0, \\
d_{45} &= 0, \\
d_{46} &= 0, \\
d_{43} &= 0.00694\lambda_1^n + 0.00132\lambda_2^n + 1.4\lambda_4^n - 2.49 \times 10^{-16}\lambda_5^n - 1.1 \times 10^{-23}\lambda_6^n + 6.15 \times 10^{-19}\lambda_7^n - 1.41, \\
d_{44} &= 2.12\lambda_1^n - 0.00132\lambda_2^n - 1.4\lambda_4^n + 3.24 \times 10^{-16}\lambda_5^n + 1.16 \times 10^{-23}\lambda_6^n - 7.96 \times 10^{-19}\lambda_7^n + 0.281, \\
d_{47} &= -111.0\lambda_1^n - 0.00503\lambda_2^n - 103.0\lambda_4^n - 13.5\lambda_5^n + 3.4 \times 10^{-7}\lambda_6^n + 0.0322\lambda_7^n + 1133.0, \\
d_{51} &= 0, \\
d_{52} &= 0, \\
d_{53} &= 0, \\
d_{54} &= 0, \\
d_{56} &= 0, \\
d_{57} &= 0, \\
d_{55} &= 4.62\lambda_1^n - 0.0177\lambda_2^n - 3.84 \times 10^{-4}\lambda_4^n - 1.36\lambda_5^n - 2.24, \\
d_{61} &= 0, \\
d_{62} &= 0, \\
d_{63} &= 0, \\
d_{64} &= 0, \\
d_{65} &= 0, \\
d_{67} &= 0, \\
d_{66} &= -1.64 \times 10^7\lambda_1^n - 3.31 \times 10^4\lambda_2^n - 5.95 \times 10^6\lambda_4^n - 421.0\lambda_5^n + 3.51 \times 10^4\lambda_6^n + 2.23 \times 10^5, \\
\end{align*} \]
with the eigenvalues of matrix $D^{-1}$ and $D$ as $\lambda_1 = 1.016$, $\lambda_2 = 3.1253$, $\lambda_3 = 1$, $\lambda_4 = 0.96$, $\lambda_5 = 1.1765$, $\lambda_6 = 14.2857$, and $\lambda_7 = 1.0526$ and $\lambda_u = 0.9843$, $\lambda_v = 0.32$, $\lambda_c = 1$, $\lambda_d = 1.0417$, $\lambda_e = 0.85$, $\lambda_f = 0.07$, and $\lambda_a = 0.95$, respectively.

From the calculation process of solution (63), the effect of isolated discrete time domains on model (34) is obvious. When $T = Z$, we get

$$a_{i,t} = a_{i,0} \left( 1 + \sum_{i=1}^{7} \frac{d_{i,i}M_i(t)}{t} \right),$$

$$c_{i,t} = c_{i,0} \left( 1 + \sum_{i=1}^{7} \frac{d_{j,j}M_j(t)}{t} \right).$$

(72)

It is worth noting that system (72) is a closed form optimal solution for model (34). System (72) explains the distribution of wealth among agents as the trading time interval is uniform. The risk aversion parameter $\delta_t$ depicts the influence of agent’s risk aversion on wealth distribution. $\phi$ and $\varphi$ describe the distribution of wealth among agents as the agents predict market risks.

6. Conclusion

In this paper, a wealth distribution model on isolated discrete time domains is given to describe the process of wealth distribution. We assume that agents have different degrees of risk aversion. The HARA utility function is employed to characterize the degrees of risk aversion of agents, including decreasing relative risk aversion, increasing relative risk aversion, and constant relative risk aversion. We consider the influence of agent’s expectation and average consumption level on wealth distribution. In addition, we allow agents to adopt certain trading strategies to maximize their utility and improve their wealth status. The Euler equation and transversality condition of the nonlinear stochastic dynamic model on isolated discrete time domains are given to prove the existence of the optimal solution of the model. The optimal solution of the wealth distribution model is obtained by using the method of solving the rational expectation model on isolated discrete time domains. A numerical example is given to illustrate the advantages of the wealth distribution model on isolated discrete time domains. The main advantage of the wealth distribution model is that, unlike the wealth distribution models in which wealth exchange is assumed to take place at evenly spaced intervals, the intervals between wealth exchange are of any arbitrary length.

Appendix

A. The Derivation of Equation (44)

Log-linearizing $y_{i,\sigma(t)} / (2\lambda + \lambda r_{\sigma(t)} - 1)$, we get

$$\frac{1}{2\lambda + \lambda r_{\sigma(t)} - 1} y_{i,\sigma(t)} \approx \frac{1}{2\lambda + \lambda r_{\sigma(0)} - 1} y_{i,\sigma(0)} + \frac{\lambda y_{i,\sigma(0)} r_{\sigma(0)} - r_{\sigma(t)} - r_{\sigma(0)}}{2\lambda + \lambda r_{\sigma(0)} - 1}$$

$$\approx \frac{1}{2\lambda + \lambda r_{\sigma(0)} - 1} y_{i,\sigma(0)} + \frac{\lambda y_{i,\sigma(0)} r_{\sigma(0)} - r_{\sigma(t)} - r_{\sigma(0)}}{2\lambda + \lambda r_{\sigma(0)} - 1} y_{i,\sigma(t)}$$

$$\approx \frac{1}{2\lambda + \lambda r_{\sigma(0)} - 1} y_{i,\sigma(0)} + \frac{\lambda y_{i,\sigma(0)} r_{\sigma(0)} - r_{\sigma(t)} - r_{\sigma(0)}}{2\lambda + \lambda r_{\sigma(0)} - 1} y_{i,\sigma(t)}$$

$$\approx \frac{1}{2\lambda + \lambda r_{\sigma(0)} - 1} y_{i,\sigma(0)} + \frac{\lambda y_{i,\sigma(0)} r_{\sigma(0)} - r_{\sigma(t)} - r_{\sigma(0)}}{2\lambda + \lambda r_{\sigma(0)} - 1} y_{i,\sigma(t)}$$

$$\approx \frac{1}{2\lambda + \lambda r_{\sigma(0)} - 1} y_{i,\sigma(0)} + \frac{\lambda y_{i,\sigma(0)} r_{\sigma(0)} - r_{\sigma(t)} - r_{\sigma(0)}}{2\lambda + \lambda r_{\sigma(0)} - 1} y_{i,\sigma(t)}$$

(71)
Similarly, log-linearizing the first and second terms on the right side of equation (41a), we obtain

\[ a_{i,t} \approx \frac{1}{2\lambda + \lambda r_{\sigma(t)}} a_{i,\sigma(t)} \]

\[ = a_{i,0}(1 + \tilde{a}_{i,t}) - \frac{a_{i,0}}{2\lambda + \lambda r_{\sigma(0)} - 1} - \frac{1}{2\lambda + \lambda r_{\sigma(0)} - 1} (a_{i,\sigma(t)} - a_{i,\sigma(0)}) + \frac{\lambda a_{i,\sigma(0)} r_{\sigma(0)}}{(2\lambda + \lambda r_{\sigma(0)} - 1)^2} (r_{\sigma(t)} - r_{\sigma(0)}). \]  \hspace{1cm} (A.2)

Thus, we have

\[ a_{i,t} \approx \frac{1}{2\lambda + \lambda r_{\sigma(t)}} a_{i,\sigma(t)} \]

\[ = a_{i,0}(1 + \tilde{a}_{i,t}) - \frac{a_{i,0}}{2\lambda + \lambda r_{\sigma(0)} - 1} - \frac{a_{i,\sigma(t)} - a_{i,\sigma(0)}}{2\lambda + \lambda r_{\sigma(0)} - 1} + \frac{\lambda a_{i,\sigma(0)} r_{\sigma(0)}}{(2\lambda + \lambda r_{\sigma(0)} - 1)^2} r_{\sigma(t)} - r_{\sigma(0)}. \]  \hspace{1cm} (A.3)

Combining equations (A.1) and (A.3), we get

\[ c_{i,t} \approx c_{i,0} + c_{i,0} \tilde{c}_{i,t} \]

\[ = a_{i,0}(1 + \tilde{a}_{i,t}) - \frac{a_{i,0}}{2\lambda + \lambda r_{\sigma(0)} - 1} - \frac{a_{i,\sigma(0)} \tilde{a}_{i,\sigma(t)}}{2\lambda + \lambda r_{\sigma(0)} - 1} + \frac{\lambda a_{i,\sigma(0)} r_{\sigma(0)}}{(2\lambda + \lambda r_{\sigma(0)} - 1)^2} r_{\sigma(t)} + \frac{y_{i,\sigma(0)}}{2\lambda + \lambda r_{\sigma(0)} - 1} \tilde{y}_{i,\sigma(t)} \]

\[ + \frac{\lambda \tilde{y}_{i,\sigma(t)}}{2\lambda + \lambda r_{\sigma(0)} - 1} (a_{i,\sigma(0)} \tilde{r}_{\sigma(t)} - y_{i,\sigma(0)} \tilde{r}_{\sigma(t)}). \]  \hspace{1cm} (A.4)

Based on initial state (42a), equation (A.4) becomes

\[ c_{i,t} \approx c_{i,0} + c_{i,0} \tilde{c}_{i,t} + \frac{y_{i,\sigma(0)} \tilde{y}_{i,\sigma(t)} - a_{i,\sigma(0)} \tilde{a}_{i,\sigma(t)}}{2\lambda + \lambda r_{\sigma(0)} - 1} \]

\[ + \frac{\lambda r_{\sigma(0)} \tilde{r}_{\sigma(t)}}{(2\lambda + \lambda r_{\sigma(0)} - 1)^2} (a_{i,\sigma(0)} - y_{i,\sigma(0)}) \tilde{r}_{\sigma(t)}. \]  \hspace{1cm} (A.5)

Thus, we obtain

\[ c_{i,0} \tilde{c}_{i,t} \approx a_{i,0} \tilde{a}_{i,t} + \frac{y_{i,\sigma(0)} \tilde{y}_{i,\sigma(t)} - a_{i,\sigma(0)} \tilde{a}_{i,\sigma(t)}}{2\lambda + \lambda r_{\sigma(0)} - 1} \]

\[ + \frac{\lambda r_{\sigma(0)} \tilde{r}_{\sigma(t)}}{(2\lambda + \lambda r_{\sigma(0)} - 1)^2} (a_{i,\sigma(0)} - y_{i,\sigma(0)}) \tilde{r}_{\sigma(t)}. \]  \hspace{1cm} (A.6)

Since \( k_t = 2\lambda + \lambda r_t - 1 \), equation (A.6) is rewritten as

\[ c_{i,0} \tilde{c}_{i,t} \approx a_{i,0} \tilde{a}_{i,t} + \frac{1}{k_{\sigma(0)}} (y_{i,\sigma(0)} \tilde{y}_{i,\sigma(t)} - a_{i,\sigma(0)} \tilde{a}_{i,\sigma(t)}) \]

\[ + \frac{\lambda r_{\sigma(0)} \tilde{r}_{\sigma(t)}}{k_{\sigma(0)}^2} (a_{i,\sigma(0)} - y_{i,\sigma(0)}) \tilde{r}_{\sigma(t)}. \]  \hspace{1cm} (A.7)

Dividing both sides of equation (A.7) by \( c_{i,0} \), we get equation (44).

**B. The Derivation of Equation (48)**

For system (39), we log-linearize the first equation. That is,
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\begin{align*}
\beta_t \frac{\mu(t)}{\mu_t} k_{\tau(t)} &= \beta_t \frac{\mu_0}{\mu_0} \left( \frac{c_{i,t} - \gamma \bar{F}_t}{c_{i,\sigma(t)} - \gamma \bar{F}_{\sigma(t)}} \right)^{\delta_t} k_{\tau(t)} \\
&\approx \beta_t \frac{\mu_0}{\mu_0} \left( \frac{c_{i,\sigma(t)} - \gamma \bar{F}_{\sigma(t)}}{c_{i,\sigma(t)} - \gamma \bar{F}_{\sigma(t)}} \right) k_{\tau(t)} \\
&\quad + \beta_t \frac{c_{i,\sigma(t)} - \gamma \bar{F}_{\sigma(t)}}{c_{i,\sigma(t)} - \gamma \bar{F}_{\sigma(t)}} \left( \mu_t - \mu_0 \right) k_{\tau(\sigma(t))} \\
- \beta_t \frac{c_{i,0} - \gamma \bar{F}_0}{\mu_0} \left( \frac{c_{i,0} - \gamma \bar{F}_0}{c_{i,\sigma(t)} - \gamma \bar{F}_{\sigma(t)}} \right)^{\delta_t} k_{\tau(\sigma(t))} \\
&\quad + \beta_t \frac{c_{i,0} - \gamma \bar{F}_0}{\mu_0} \left( c_{i,\sigma(t)} - \gamma \bar{F}_{\sigma(t)} \right)^{\delta_t} k_{\tau(\sigma(t))} \left( c_{i,t} - c_{i,0} \right) \\
- \gamma \beta_t \frac{c_{i,0} - \gamma \bar{F}_0}{\mu_0} \left( c_{i,\sigma(t)} - \gamma \bar{F}_{\sigma(t)} \right)^{\delta_t} k_{\tau(\sigma(t))} \\
&\quad - \gamma \beta_t \frac{c_{i,0} - \gamma \bar{F}_0}{\mu_0} \left( c_{i,\sigma(t)} - \gamma \bar{F}_{\sigma(t)} \right)^{\delta_t} k_{\tau(\sigma(t))} \\
&\quad + \frac{\mu_0}{\mu_0} \left( c_{i,\sigma(t)} - \gamma \bar{F}_{\sigma(t)} \right)^{\delta_t} k_{\tau(\sigma(t))} \\
&\quad + \frac{\mu_0}{\mu_0} \left( c_{i,0} - \gamma \bar{F}_0 \right)^{\delta_t} k_{\tau(\sigma(t))} - \delta \left( c_{i,0} - \gamma \bar{F}_0 \right)^{\delta_t} k_{\tau(\sigma(t))} \\
&\quad + \frac{\mu_0}{\mu_0} \left( c_{i,0} - \gamma \bar{F}_0 \right)^{\delta_t} k_{\tau(\sigma(t))} \left( c_{i,\tau(t)} - c_{i,0} \right). \\
\end{align*}

(B.1)

Since \( \hat{\chi}_t = (x_t - x_0)/\lambda_0 \), equation (B.1) is turned into

\begin{align*}
\beta_t \frac{\mu(t)}{\mu_t} \left( c_{i,t} - \gamma \bar{F}_t \right)^{\delta_t} k_{\tau(\sigma(t))} \\
&\approx \beta_t \frac{\mu_0}{\mu_0} \left( c_{i,\sigma(t)} - \gamma \bar{F}_{\sigma(t)} \right)^{\delta_t} k_{\tau(\sigma(t))} \\
&\quad + \beta_t \frac{\mu_0}{\mu_0} \left( c_{i,\sigma(t)} - \gamma \bar{F}_{\sigma(t)} \right)^{\delta_t} k_{\tau(\sigma(t))} \left( c_{i,t} - c_{i,0} \right) \\
- \beta_t \frac{\mu_0}{\mu_0} \left( c_{i,0} - \gamma \bar{F}_0 \right)^{\delta_t} k_{\tau(\sigma(t))} \\
&\quad + \beta_t \frac{\mu_0}{\mu_0} \left( c_{i,0} - \gamma \bar{F}_0 \right)^{\delta_t} k_{\tau(\sigma(t))} \left( c_{i,\tau(t)} - c_{i,0} \right). \\
\end{align*}

(B.2)

By elementary algebra, we get

\begin{align*}
\beta_t \frac{\mu(t)}{\mu_t} \left( c_{i,t} - \gamma \bar{F}_t \right)^{\delta_t} k_{\tau(\sigma(t))} \\
&\approx \beta_t \frac{\mu_0}{\mu_0} \left( c_{i,0} - \gamma \bar{F}_0 \right)^{\delta_t} k_{\tau(\sigma(t))} \left[ 1 + \mu_{\sigma(t)} - \mu_t \right] \\
&\quad + \beta_t \frac{c_{i,0} - \gamma \bar{F}_0}{c_{i,\sigma(t)} - \gamma \bar{F}_{\sigma(t)}} \left( c_{i,0} - \gamma \bar{F}_0 \right)^{\delta_t} k_{\tau(\sigma(t))} \\
&\quad + \frac{\delta t_{\gamma \bar{F}_0} / \lambda_0}{c_{i,0} - \gamma \bar{F}_0} \left( c_{i,0} - \gamma \bar{F}_0 \right)^{\delta_t} k_{\tau(\sigma(t))} \\
&\quad + \frac{\delta t_{\gamma \bar{F}_0} / \lambda_0}{c_{i,0} - \gamma \bar{F}_0} \left( c_{i,0} - \gamma \bar{F}_0 \right)^{\delta_t} k_{\tau(\sigma(t))}. \\
\end{align*}

(B.3)

On the basis of equations (42d) and (B.3), we get equation (48). Similarly, we log-linearize the second equation in system (39) and obtain equation (49).

C. The Derivation of Equation (50)

To log-linearize equation (41d), we rewrite it at \( \sigma(t) \) and obtain

\[ \log y_{\sigma(t)} = (1 - \phi) \log y_{1,0} + \phi \log y_{1,0} + \phi_1 \log y_{1,0} + \epsilon_{\sigma(t)}. \]  

(C.1)

Log-linearizing on both sides of equation (C.1), we get

\[ \log y_{\sigma(t)} + \frac{1}{y_{\sigma(t)}} (y_{\sigma(t)} - y_{\sigma(t)}) = (1 - \phi) \log y_{1,0} + \phi_1 \log y_{1,0} + \epsilon_{\sigma(t)}. \]  

(C.2)

\[ \log y_{\sigma(t)} + \frac{1}{y_{\sigma(t)}} (y_{\sigma(t)} - y_{\sigma(t)}) = (1 - \phi) \log y_{1,0} + \phi_1 \log y_{1,0} + \epsilon_{\sigma(t)}. \]  

(C.3)

Since \( \log y_{\sigma(t)} = (1 - \phi) \log y_{1,0} + \phi_1 \log y_{1,0} + \epsilon_{\sigma(t)} \), we obtain equation (50).

D. The Derivation of Equation (54)

To obtain the log-linearization expression of equation (41f), we write equation (41f) as follows:

\[ \log r_{\sigma(t)} = \varphi \log r_0 + (1 - \varphi) \log r_t + \eta_{\sigma(t)}. \]  

(D.1)

Log-linearizing equation (D.1), we obtain

\[ \log r_{\sigma(t)} + \frac{1}{r_{\sigma(t)}} (r_{\sigma(t)} - r_{\sigma(t)}) = \varphi \log r_0 + (1 - \varphi) \log r_0 + (1 - \varphi) \frac{1}{r_0} (r_t - r_0) + \eta_{\sigma(t)}. \]  

(D.2)

\[ \log r_{\sigma(t)} + \frac{1}{r_{\sigma(t)}} (r_{\sigma(t)} - r_{\sigma(t)}) = \varphi \log r_0 + (1 - \varphi) \log r_0 + (1 - \varphi) \frac{1}{r_0} (r_t - r_0) + \eta_{\sigma(t)}. \]  

(D.3)
Since \( \log r_{\phi(0)} = \varphi \log r_0 + (1 - \varphi) \log r_0 + \eta e^{\phi(0)} \), equation (D.3) becomes equation (54).

**Data Availability**

No data were used to support this study.

**Conflicts of Interest**

No potential conflicts of interest have been reported by the authors.

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