The (Un)expected Effects of Applying Standard Cleansing Models to Human Ratings on Compositionality

Stephen Roller†‡ Sabine Schulte im Walde † Silke Scheible †

†Department of Computer Science
The University of Texas at Austin
roller@cs.utexas.edu
‡Institut für Maschinelle Sprachverarbeitung
Universität Stuttgart
{schulte,scheible}@ims.uni-stuttgart.de

Abstract

Human ratings are an important source for evaluating computational models that predict compositionality, but like many data sets of human semantic judgements, are often fraught with uncertainty and noise. However, despite their importance, to our knowledge there has been no extensive look at the effects of cleansing methods on human rating data. This paper assesses two standard cleansing approaches on two sets of compositionality ratings for German noun-noun compounds, in their ability to produce compositionality ratings of higher consistency, while reducing data quantity. We find (i) that our ratings are highly robust against aggressive filtering; (ii) Z-score filtering fails to detect unreliable item ratings; and (iii) Minimum Subject Agreement is highly effective at detecting unreliable subjects.

1 Introduction

Compounds have long been a reoccurring focus of attention within theoretical, cognitive, and computational linguistics. Recent manifestations of interest in compounds include the Handbook of Compounding (Lieber and Stekauer, 2009) on theoretical perspectives, and a series of workshops and special journal issues with respect to the computational perspective (Journal of Computer Speech and Language, 2005; Language Resources and Evaluation, 2010; ACM Transactions on Speech and Language Processing, to appear). Some work has focused on modeling meaning and compositionality for specific classes, such as particle verbs (McCarthy et al., 2003; Bannard, 2005; Cook and Stevenson, 2006); adjective-noun combinations (Baroni and Zamparelli, 2010; Boleda et al., 2013); and noun-noun compounds (Reddy et al., 2011b; Reddy et al., 2011a). Others have aimed at predicting the compositionality of phrases and sentences of arbitrary type and length, either by focusing on the learning approach (Socher et al., 2011); by integrating symbolic models into distributional models (Coecke et al., 2011; Grefenstette et al., 2013); or by exploring the arithmetic operations to predict compositionality by the meaning of the parts (Widdows, 2008; Mitchell and Lapata, 2010).

An important resource in evaluating compositionality has been human compositionality ratings, in which human subjects are asked to rate the degree to which a compound is transparent or opaque. Transparent compounds, such as raincoat, have a meaning which is an obvious combination of its constituents, e.g., a raincoat is a coat against the rain. Opaque compounds, such as hot dog, have little or no relation to one or more of their constituents: a hot dog need not be hot, nor is it (hopefully) made of dog. Other words, such as ladybug, are transparent with respect to just one constituent. As many words do not fall clearly into one category or the other, subjects are typically asked to rate the compositionality of words or phrases on a scale, and the mean of several judgements is taken as the gold standard.

Like many data sets of human judgements, compositionality ratings can be fraught with large quantities of uncertainty and noise. For example, participants typically agree on items that are clearly transparent or opaque, but will often disagree about the
gray areas in between. Such uncertainty represents an inherent part of the semantic task and is the major reason for using the mean ratings of many subjects.

Other types of noise, however, are undesirable, and should be eliminated. In particular, we wish to examine two types of potential noise in our data. The first type of noise (Type I noise: uncertainty), comes from when a subject is unfamiliar or uncertain about particular words, resulting in sporadically poor judgements. The second type of noise (Type II noise: unreliability), occurs when a subject is consistently unreliable or uncooperative. This may happen if the subject misunderstands the task, or if a subject simply wishes to complete the task as quickly as possible. Judgements collected via crowdsourcing are especially prone to this second kind of noise, when compared to traditional pen-and-paper experiments, since participants aim to maximize their hourly wage.\(^2\)

In this paper, we apply two standard cleansing methods (Ben-Gal, 2005; Maletic and Marcus, 2010), that have been used on similar rating data before (Reddy et al., 2011b), on two data sets of compositionality ratings of German noun-noun compounds. We aim to address two main points. The first is to assess the cleansing approaches in their ability to produce compositionality ratings of higher quality and consistency, while facing a reduction of data mass in the cleansing process. In particular, we look at the effects of removing outlier judgements resulting from uncertainty (Type I noise) and dropping unreliable subjects (Type II noise). The second issue is to assess the overall reliability of our two rating data sets: Are they clean enough to be used as gold standard models in computational linguistics approaches?

2 Compositionality Ratings

Our focus of interest is on German noun-noun compounds (see Fleischer and Barz (2012) for a detailed overview), such as Ahornblatt ‘maple leaf’ and Feuerwerk ‘fireworks’, and Obstkuchen ‘fruit cake’ where both the head and the modifier are nouns. We rely on a subset of 244 noun-noun compounds collected by von der Heide and Borgwaldt (2009), who created a set of 450 concrete, depictable German noun compounds according to four compositionality classes (transparent+transparent, transparent+opaque, opaque+transparent, opaque+opaque).

We are interested in the degrees of compositionality of the German noun-noun compounds, i.e., the relation between the meaning of the whole compound (e.g., Feuerwerk) and the meaning of its constituents (e.g., Feuer ‘fire’ and Werk ‘opus’). We work with two data sets of compositionality ratings for the compounds. The first data set, the individual compositionality ratings, consists of participants rating the compositionality of a compound with respect to each of the individual constituents. These judgements were collected within a traditional controlled, pen-and-paper setting. For each compound-constituent pair, 30 native German speakers rated the compositionality of the compound with respect to its constituent on a scale from 1 (opaque/non-compositional) to 7 (transparent/compositional). The subjects were allowed to omit ratings for unfamiliar words, but very few did; of the 14,640 possible ratings judgements, only 111 were left blank. Table 1 gives several examples of such ratings. We can see that Fliegenpilz ‘toadstool’ is an example of a very opaque (non-compositional) word with respect to Fliege ‘housefly/bow tie’; it has little to do with either houseflies or bow ties. On the other hand Teetasse ‘cup’ intended for Tee ‘tea’.

The second data set, the whole compositionality ratings consists of participants giving a single rating for the entire compound. These ratings, previously unpublished, reflect a very different view of the same compounds. Rather than rating compounds with respect to their constituents, subjects were asked to give a single rating for the entire compound using the same 1-7 scale as before. The ratings were collected via Amazon Mechanical Turk (AMT). The data was controlled for spammers by removing subjects who failed to identify a number of fake words. Subjects who rated less than 10 compounds or had a low AMT reputation were also removed. The resulting data represents 150 different subjects with roughly 30 ratings per compound. Most participants rated only a few dozen items. We can see examples of these ratings in Table 2.

\(^2\)See Callison-Burch and Dredze (2010) for a collection of papers on data collected with AMT. While the individual approaches deal with noise in individual ways, there is no general approach to clean crowdsourcing data.
Table 1: Sample compositionality ratings for three compounds with respect to their constituents. We list the mean rating for only these 4 subjects to facilitate examples. The Combined column is the geometric mean of both constituents.

| Compound W.R.T. | Subject 1 | Subject 2 | Subject 3 | Subject 4 | Mean | Comb. |
|-----------------|-----------|-----------|-----------|-----------|------|-------|
| Fliegenpilz ‘toadstool’ | Fliege ‘housefly/bow tie’ | 3 | 1 | 1 | 2 | 1.75 | 3.37 |
| Fliegenpilz ‘toadstool’ | Pilz ‘mushroom’ | 5 | 7 | 7 | 7 | 6.50 | 3.37 |
| Sonnenblume ‘sunflower’ | Sonne ‘sun’ | 4 | 3 | 1 | 2 | 2.50 | 4.11 |
| Sonnenblume ‘sunflower’ | Blume ‘flower’ | 7 | 7 | 7 | 6 | 6.75 | 4.11 |
| Teetasse ‘teacup’ | Tee ‘tea’ | 6 | 6 | 4 | 2 | 4.50 | 4.50 |
| Teetasse ‘teacup’ | Tasse ‘cup’ | 7 | 6 | 4 | 1 | 4.50 | 4.50 |

Table 2: Example whole compositionality ratings for three compounds. Note that Subject 1 chose not to rate Fliegenpilz, so the mean is computed using only the three available judgements.

| Compound | Subject 1 | Subject 2 | Subject 3 | Subject 4 | Mean |
|----------|-----------|-----------|-----------|-----------|------|
| Fliegenpilz ‘toadstool’ | - | 2 | 1 | 2 | 2.67 |
| Sonnenblume ‘sunflower’ | 3 | 3 | 1 | 2 | 2.75 |
| Teetasse ‘teacup’ | 7 | 7 | 7 | 6 | 6.75 |

3 Methodology

In order to check on the reliability of compositionality judgements in general terms as well as with regard to our two specific collections, we applied two standard cleansing approaches\(^3\) to our rating data: Z-score filtering is a method for filtering Type I noise, such as random guesses made by individuals when a word is unfamiliar. Minimum Subject Agreement is a method for filtering out Type II noise, such as subjects who seem to misunderstand the rating task or rarely agree with the rest of the population. We then evaluated the original vs. cleaned data by one intrinsic and one extrinsic task. Section 3.1 presents the two evaluations and the unadulterated, baseline measures for our experiments. Sections 3.2.1 and 3.2.2 describe the cleansing experiments and results.

3.1 Evaluations and Baselines

For evaluating the cleansing methods, we propose two metrics, an intrinsic and an extrinsic measure.

3.1.1 Intrinsic Evaluation:
Consistency between Rating Data Sets

The intrinsic evaluation measures the consistency between our two ratings sets individual and whole. Assuming that the compositionality ratings for a compound depend heavily on both constituents, we expect a strong correlation between the two data sets. For a compound to be rated transparent as a whole, it should be transparent with respect to both of its constituents. Compounds which are highly transparent with respect to only one of their constituents should be penalized appropriately.

In order to compute a correlation between the whole ratings (which consist of one average rating per compound) and the individual ratings (which consist of two average ratings per compound, one for each constituent), we need to combine the individual ratings to arrive at a single value. We use the geometric mean to combine the ratings, which is effectively identical to the multiplicative methods in Widows (2008), Mitchell and Lapata (2010) and Reddy et al. (2011b).\(^4\) For example, using our means listed in Table 1, we may compute the combined rating for Sonnenblume as \(\sqrt{6.75 \times 2.50} \approx 4.11\). These combined ratings are computed for all compounds, as listed in the “Comb.” column of Table 1. We then compute our consistency measure as the Spearman’s \(\rho\) rank correlation between these combined individual ratings with the whole ratings (“Mean” in Table 2). The original, unadulterated data sets have a consistency measure of 0.786, indicating that, despite the very different collection methodologies, the two ratings sets largely agree.

3.1.2 Extrinsic Evaluation:
Correlation with Association Norms

The extrinsic evaluation compares the consistency

\(^3\)See Ben-Gal (2005) or Maletic and Marcus (2010) for overviews of standard cleansing approaches.

\(^4\)We also tried the arithmetic mean, but the multiplicative method always performs better.
between our two rating sets individual and whole with evidence from a large collection of association norms. Association norms have a long tradition in psycholinguistic research to investigate semantic memory, making use of the implicit notion that associates reflect meaning components of words (Deese, 1965; Miller, 1969; Clark, 1971; Nelson et al., 1998; Nelson et al., 2000; McNamara, 2005; de Deyne and Storms, 2008). They are collected by presenting a stimulus word to a subject and collecting the first words that come to mind.

We rely on association norms that were collected for our compounds and constituents via both a large scale web experiment and Amazon Mechanical Turk (Schulte im Walde et al., 2012) (unpublished). The resulting combined data set contains 85,049/34,560 stimulus-association tokens/types for the compound and constituent stimuli. Table 3 gives examples of associations from the data set for some stimuli.

The guiding intuition behind comparing our rating data sets with association norms is that a compound which is compositional with respect to a constituent should have similar associations as its constituent (Schulte im Walde et al., 2012).

To measure the correlation of the rating data with the association norms, we first compute the Jaccard similarity that measures the overlap in two sets, ranging from 0 (perfectly dissimilar) to 1 (perfectly similar). The Jaccard is defined for two sets, $A$ and $B$, as

$$J(A, B) = \frac{|A \cap B|}{|A \cup B|}.$$

For example, we can use Table 3 to compute the Jaccard similarity between $\text{Sonnenblume}$ and $\text{Sonne}$:

$$\frac{|\{\text{Sommer}\}|}{|\{\text{gelb}, \text{Sommer}, \text{Kerne}, \text{warm}, \text{hell}\}|} = 0.20.$$

After computing the Jaccard similarity between all compounds and constituents across the association norms, we correlate this association overlap with the average individual ratings (i.e., column “Mean” in Table 1) using Spearman’s $\rho$. This correlation “Assoc Norm (Indiv)” reaches $\rho = 0.638$ for our original data. We also compute a combined Jaccard similarity using the geometric mean, e.g.

$$\sqrt{J(\text{Fliegenpilz}, \text{Fliege}) \cdot J(\text{Fliegenpilz}, \text{Pilz})},$$

and calculate Spearman’s $\rho$ with the whole ratings (i.e., column “Mean” in Table 2). This correlation “Assoc Norm (Whole)” reaches $\rho = 0.469$ for our original data.

### 3.2 Data Cleansing

We applied the two standard cleansing approaches, Z-score Filtering and Minimum Subject Agreement, to our rating data, and evaluated the results.

#### 3.2.1 Z-score Filtering

Z-score filtering is a method to filter out Type I noise, such as random guesses made by individuals when a word is unfamiliar. It makes the simple assumption that each item’s ratings should be roughly normally distributed around the “true” rating of the item, and throws out all outliers which are more than $z^*$ standard deviations from the item’s mean. With regard to our compositionality ratings, for each item $i$ (i.e., a compound in the whole data, or a compound-constituent pair in the individual data) we compute the mean $\bar{x}_i$ and standard deviation $\sigma_i$ of the ratings for the given item. We then remove all values from $x_i$ where

$$|x_i - \bar{x}_i| > \sigma_i z^*,$$

with the parameter $z^*$ indicating the maximum allowed Z-score of the item’s ratings. For example, if a particular item has ratings of $x_i = (1, 2, 1, 6, 1, 1)$, then the mean $\bar{x}_i = 2$ and the standard deviation

| Word   | Example Associations |
|--------|----------------------|
| Fliegenpilz  | 'toadstool' giftig 'poisonous', rot 'red', Wald 'forest' |
| Fliege  | 'housely/bow tie' nervig 'annoying', summen 'to buzz', Insekt 'insect' |
| Pilz   | 'mushroom' Wald 'forest', giftig 'poisonous', sammeln 'to gather' |
| Sonnenblume  | 'sunflower' gelb 'yellow', Sommer 'summer', Kerne 'seeds' |
| Sonne  | 'sun' Sommer 'summer', warm 'warm', hell 'bright' |
| Blume  | 'flower' Wiese 'meadow', Daft 'smell', Rose 'rose' |

Table 3: Example association norms for two German compounds and their constituents.
Figure 1: Intrinsic and Extrinsic evaluation of Z-score filtering. We see that Z-score filtering makes a minimal difference when filtering is strict, and is slightly detrimental with more aggressive filtering.

Figures 1a and 1b show the results for the intrinsic evaluation of Z-score filtering. The solid black line represents the consistency of the filtered individual ratings with the unadulterated whole ratings. The dotted orange line shows the consistency of the filtered whole ratings with the unadulterated individual ratings, and the dashed purple line shows the consistency between the data sets when both are filtered. In comparison, the consistency between the unadulterated data sets is provided by the horizontal gray line. We see that Z-score filtering overall has a minimal effect on the consistency of the two data sets. It provides very small improvements with high Z-scores, but is slightly detrimental at more aggressive levels.

Figure 1b shows the effects of Z-score filtering with our extrinsic evaluation of correlation with association norms. At all levels of filtering, we see that correlation with association norms remains mostly independent of the level of filtering.

An important factor to consider when evaluating these results is the amount of data dropped at each of the filtering levels. Figure 2 shows the data retention rate for the different data sets and levels. As expected, more aggressive filtering results in a substantially lower data retention rate. Comparing this curve to the consistency ratings gives a clear picture: the decrease in consistency is probably mostly due to the decrease in available data but not due to filtering outliers. As such, we believe that Z-score filtering does not substantially improve data quality, but may be safely applied with a conservative maximum allowed Z-score.

Figure 2: The data retention rate of Z-score filtering. Data retention drops rapidly with aggressive filtering.

Filtering Outliers  Figure 1a shows the results for the intrinsic evaluation of Z-score filtering. The solid black line represents the consistency of the filtered individual ratings with the unadulterated whole ratings. The dotted orange line shows the consistency of the filtered whole ratings with the unadulterated individual ratings, and the dashed purple line shows the consistency between the data sets when both are filtered. In comparison, the consistency between the unadulterated data sets is provided by the horizontal gray line. We see that Z-score filtering overall has a minimal effect on the consistency of

Filtering Artificial Noise  Z-score filtering has little impact on the consistency of the data, but we would like to determine whether this is due because our data being very clean, so the filtering does not apply, or Z-score filtering not being able to detect the Type I noise. To test these two possibilities, we artificially introduce noise into our data sets: we create 100 variations of the original ratings matrices, where with 0.25 probability, each entry in the matrix was
Figure 3: Ability of Z-score filtering at removing artificial noise added in the (a) individual and (b) whole judgements. The orange lines represent the consistency of the data with the noise, but no filtering, while the black lines indicate the consistency after Z-score filtering. Z-score filtering appears to be unable to find uniform random noise in either situation.

Figure 4: Intrinsic and Extrinsic evaluation of Minimum Subject Agreement filtering. We see virtually no gains using subject filtering, and the individual judgements are quite hindered by aggressive filtering.
replaced with a uniform random integer between 1 and 7. That is, roughly 1 in 4 of the entries in the original matrix were replaced with random, uniform noise. We then apply Z-score filtering on each of these noisy matrices and report their average consistency with its companion, unadulterated matrix. That is, we add noise to the individual ratings matrix, and then compare its consistency with the original whole ratings matrix, and vice versa. Thus if we are able to detect and remove the artificial noise, we should see higher consistencies in the filtered matrix over the noisy matrix.

Figure 3 shows the results of adding noise to the original data sets. The lines indicate the averages over all 100 matrix variations, while the shaded areas represent the 95% confidence intervals. Surprisingly, even though 1/4 entries in the matrix were replaced with random values, the decrease in consistency is relatively low in both settings. This likely indicates our data already has high variance. Furthermore, in both settings, we do not see any increase in consistency from Z-score filtering. We must conclude that Z-score appears ineffective at removing Type I noise in compositionality ratings.

We also tried introducing artificial noise in a second way, where judgements were not replaced with a uniformly random value, but a fixed offset of either +3 or -3, e.g., 4’s became either 1’s or 7’s. Again, the values were changed with probability of 0.25. The results were remarkably similar, so we do not include them here.

### 3.2.2 Minimum Subject Agreement

Minimum Subject Agreement is a method for filtering out subjects who seem to misunderstand the rating task or rarely agree with the rest of the population. For each subject in our data, we compute the average ratings for each item excluding the subject. The subject’s rank agreement with the exclusive averages is computed using Spearman’s $\rho$. We can then remove subjects whose rank agreement is below a threshold, or remove the $n$ subjects with the lowest rank agreement.

**Filtering Unreliable Subjects** Figure 4 shows the effect of subject filtering on our intrinsic and extrinsic evaluations. We can see that mandating minimum subject agreement has a strong, negative impact on the individual ratings after a certain threshold is reached, but virtually no effect on the whole ratings. When we consider the corresponding data retention curve in Figure 6, the result is not surprising: the dip in performance for the individual ratings comes with a data retention rate of roughly 25%. In this way, it’s actually surprising that it does so well: with only 25% of the original data, consistency is only 5 points lower. The effects are more dramatic in the extrinsic evaluation.

On the other hand, subject filtering has almost no effect on the whole ratings. This is not surprising, as most subjects have only rated at most a few dozen items, so removing subjects corresponds to a smaller reduction in data, as seen in Figure 6. Furthermore, the subjects with the highest deviations tend to be
the subjects who rated the fewest items since their agreement is more sensitive to small changes. As such, the subjects removed tend to be the subjects with the least influence on the data set.

Removing Artificial Subject-level Noise To test the hypothesis that minimum subject agreement filtering is effective at removing Type II noise, we introduce artificial noise at the subject level. For these experiments, we create 100 variations of our matrices where \( n \) subjects have all of their ratings replaced with random, uniform ratings. We then apply subject-level filtering where we remove the \( n \) subjects who agree least with the overall averages.

Figure 5a shows the ability of detecting Type II noise in the individual ratings. The results are unsurprising, but encouraging. We see that increasing the number of randomized subjects rapidly lowers the consistency with the whole ratings. However, the cleaned whole ratings matrix maintains a fairly high consistency, indicating that we are doing a nearly perfect job at identifying the noisy individuals.

Figure 5b shows the ability of detecting Type II noise in the whole ratings. Again, we see that the cleaned noisy ratings have a higher consistency than the noisy ratings, indicating the efficacy of subject agreement filtering at detecting unreliable subjects. The effect is less pronounced in the whole ratings than the individual ratings due to the lower proportion of subjects being randomized.

Identification of Spammers Removing subjects with the least agreement lends itself to another sort of evaluation: predicting subjects rejected during data collection. As discussed in Section 2, subjects who failed to identify the fake words or had an overall low reputability were filtered from the data before any analysis. To test the quality of minimum subject agreement, we reconstructed the data set where these previously rejected users were included, rather than removed. Subjects who rated fewer than 10 items were still excluded.

The resulting data set had a total of 242 users: 150 (62.0%) which were included in the original data, and 92 (38.0%) which were originally rejected. After constructing the modified data set, we sorted the subjects by their agreement. Of the 92 subjects with the lowest agreement, 75 of them were rejected in the original data set (81.5%). Of the 150 subjects with the highest agreement, only 17 of them were rejected from the original data set (11.3%). The typical precision-recall tradeoff obviously applies.

Curiously, we note that the minimum subject agreement at this 92nd subject was 0.457. Comparing with the curves for the individual ratings in Figures 4a and 6, we see this is the point where intrinsic consistency and data retention both begin dropping rapidly. While this may be a happy coincidence, it does seem to suggest that the ideal minimum subject agreement is roughly where the data retention rate starts rapidly turning.

Regardless, we can definitely say that minimum subject agreement is a highly effective way of rooting out spammers and unreliable participants.

4 Conclusion

In this paper, we have performed a thorough analysis of two sets of compositionality ratings to German noun-noun compounds, and assessed their reliability from several perspectives. We conclude that asking for ratings of compositionality of compound words is reasonable and that such judgements are notably reliable and robust. Even when compositionality ratings are collected in two very different settings (laboratory vs. AMT) and with different dynamics, the produced ratings are highly consistent. This is shown by the high initial correlation of the two sets of compositionality ratings. We believe this
provides strong evidence that human judgements of compositionality, or at least these particular data sets, are reasonable as gold standards for other computational linguistic tasks.

We also find that such ratings can be highly robust against large amounts of data loss, as in the case of aggressive Z-score and minimum subject agreement filtering: despite data retention rates of 10-70%, consistency between our data sets never dropped more than 6 points. In addition, we find that the correlation between compositionality ratings and association norms is substantial, but generally much lower and less sensitive than internal consistency.

We generally find Type I noise to be very difficult to detect, and Z-score filtering is mostly ineffective at eliminating unreliable item ratings. This is confirmed by both our natural and artificial experiments. At the same time, Z-score filtering seems fairly harmless at conservative levels, and probably can be safely applied in moderation with discretion.

On the other hand, we have confirmed that minimum subject agreement is highly effective at filtering out incompetent and unreliable subjects, as evidenced by both our artificial and spammer detection experiments. We conclude that, as we have defined it, Type II noise is easily detected, and removing this noise produces much higher quality data. We recommend using subject agreement as a first-pass identifier of likely unreliable subjects in need of manual review.

We would also like to explore other types of compounds, such as adjective-noun compounds (e.g. Großeltern ‘grandparents’), and compounds with more than two constituents (e.g. Bleistiftspitzmaschine ‘automatic pencil sharpener’).

Acknowledgments

We thank the SemRel group, Alexander Fraser, and the reviewers for helpful comments and feedback. The authors acknowledge the Texas Advanced Computing Center (TACC) for providing grid resources that have contributed to these results.5

References

Collin Bannard. 2005. Learning about the Meaning of Verb–Particle Constructions from Corpora. Computer Speech and Language, 19:467–478.

Marco Baroni and Roberto Zamparelli. 2010. Nouns are vectors, adjectives are matrices: Representing adjective-noun constructions in semantic space. In Proceedings of the 2010 Conference on Empirical Methods in Natural Language Processing, pages 1183–1193, Cambridge, MA, October.

Irad Ben-Gal. 2005. Outlier detection. In O. Maimon and L. Rockach, editors, Data Mining and Knowledge Discovery Handbook: A Complete Guide for Practitioners and Researchers. Kluwer Academic Publishers.

Gemma Boleda, Marco Baroni, Nghia The Pham, and Louise McNally. 2013. On adjective-noun composition in distributional semantics. In Proceedings of the 10th International Conference on Computational Semantics, Potsdam, Germany.

Chris Callison-Burch and Mark Dredze, editors. 2010. Proceedings of the NAACL/HLT Workshop on Creating Speech and Language Data with Amazon’s Mechanical Turk, Los Angeles, California.

Herbert H. Clark. 1971. Word Associations and Linguistic Theory. In John Lyons, editor, New Horizon in Linguistics, chapter 15, pages 271–286. Penguin.

Bob Coecke, Mehrnoosh Sadrzadeh, and Stephen Clark. 2011. Mathematical foundations for a compositional distributional model of meaning. Linguistic Analysis, 36(1-4):345–384.

Paul Cook and Suzanne Stevenson. 2006. Classifying Particle Semantics in English Verb-Particle Constructions. In Proceedings of the ACL/COLING Workshop on Multiword Expressions: Identifying and Exploiting Underlying Properties, Sydney, Australia.

Simon de Deyne and Gert Storms. 2008. Word associations: Norms for 1,424 dutch words in a continuous task. Behavior Research Methods, 40(1):198–205.

James Deese. 1965. The Structure of Associations in Language and Thought. The John Hopkins Press, Baltimore, MD.

Wolfgang Fleischer and Irmhild Barz. 2012. Wortbildung der deutschen Gegenwartssprache. de Gruyter.

Edward Grefenstette, G. Dinu, Y. Zhang, Mehrnoosh Sadrzadeh, and Marco Baroni. 2013. Multi-step regression learning for compositional distributional semantics. In Proceedings of the 10th International Conference on Computational Semantics, Potsdam, Germany.

Rochelle Lieber and Pavol Stekauer, editors. 2009. The Oxford Handbook of Compounding. Oxford University Press.

5http://www.tacc.utexas.edu
Jonathan I. Maletic and Adrian Marcus. 2010. Data cleansing: A prelude to knowledge discovery. In O. Maimon and L. Rokach, editors, Data Mining and Knowledge Discovery Handbook. Springer Science and Business Media, 2 edition.

Diana McCarthy, Bill Keller, and John Carroll. 2003. Detecting a Continuum of Compositionality in Phrasal Verbs. In Proceedings of the ACL-SIGLEX Workshop on Multimodal Expressions: Analysis, Acquisition and Treatment, Sapporo, Japan.

Timothy P. McNamara. 2005. Semantic Priming: Perspectives from Memory and Word Recognition. Psychology Press, New York.

George Miller. 1969. The Organization of Lexical Memory: Are Word Associations sufficient? In George A. Talland and Nancy C. Waugh, editors, The Pathology of Memory, pages 223–237. Academic Press, New York.

Jeff Mitchell and Mirella Lapata. 2010. Composition in Distributional Models of Semantics. Cognitive Science, 34:1388–1429.

Douglas L. Nelson, Cathy L. McEvoy, and Thomas A. Schreiber. 1998. The University of South Florida Word Association, Rhyme, and Word Fragment Norms.

Douglas L. Nelson, Cathy L. McEvoy, and Simon Dennis. 2000. What is Free Association and What does it Measure? Memory and Cognition, 28:887–899.

Siva Reddy, Ioannis P. Klapaftis, Diana McCarthy, and Suresh Manandhar. 2011a. Dynamic and Static Prototype Vectors for Semantic Composition. In Proceedings of the 5th International Joint Conference on Natural Language Processing, pages 705–713, Chiang Mai, Thailand.

Siva Reddy, Diana McCarthy, and Suresh Manandhar. 2011b. An Empirical Study on Compositionality in Compound Nouns. In Proceedings of the 5th International Joint Conference on Natural Language Processing, pages 210–218, Chiang Mai, Thailand.

Sabine Schulte im Walde, Susanne Borgwaldt, and Ronny Jauch. 2012. Association Norms of German Noun Compounds. In Proceedings of the 8th International Conference on Language Resources and Evaluation, pages 632–639, Istanbul, Turkey.

Richard Socher, Eric H. Huang, Jeffrey Pennington, Andrew Y. Ng, and Christopher D. Manning. 2011. Dynamic Pooling and Unfolding Recursive Autoencoders for Paraphrase Detection. In Advances in Neural Information Processing Systems 24.

Claudia von der Heide and Susanne Borgwaldt. 2009. Assoziationen zu Unter-, Basis- und Oberbegriffen. Eine explorative Studie. In Proceedings of the 9th Norddeutsches Linguistisches Kolloquium, pages 51–74.

Dominic Widdows. 2008. Semantic Vector Products: Some Initial Investigations. In Proceedings of the 2nd Conference on Quantum Interaction, Oxford, UK.