Effects of Culling on *Leptospira interrogans* Carriage by Rats

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We found that lethal, urban rat control is associated with a significant increase in the odds that surviving rats carry *Leptospira interrogans*. Our results suggest that human interventions have the potential to affect and even increase the prevalence of zoonotic pathogens within rat populations.

Norway rats (*Rattus norvegicus*) are a reservoir for *Leptospira interrogans*, the etiologic agent of the zoonotic disease leptospirosis (1). Leptospirosis affects ≈1 million persons worldwide annually and can result in kidney failure or pulmonary hemorrhage (1,2). Increasing urbanization has driven the emergence of leptospirosis in cities globally (3). Within cities, areas of poverty experience a confluence of environmental and socioeconomic factors that heighten the risk for ratborne *L. interrogans* transmission (3).

The ecology of rats and the epidemiology of *L. interrogans* within their populations are intimately connected (4). Previous research on other reservoir species suggests that anthropogenic disturbances may alter reservoir ecology, resulting in new transmission patterns (5,6). Because lethal control is a common technique used to address rat populations (7,8), we aimed to determine whether culling affects *L. interrogans* carriage by urban Norway rats.

The Study

This study, conducted in an inner-city neighborhood of Vancouver, British Columbia, Canada, during June 2016–January 2017, compared the prevalence of *L. interrogans* in rat populations before and after a kill-trapping intervention. Each study site (12 total) comprised 3 contiguous city blocks and was designated as a control site or an intervention site (Figure 1). In control sites, no kill-trapping occurred; in intervention sites, kill-trapping occurred only in the central blocks, and the 2 adjacent blocks were designated as nonkill flanking blocks. We divided trapping in each intervention site into 3 time periods: before, during, and after the intervention (Figure 2). Before and after the intervention, rats were trapped, processed, and released. During processing, rats were marked with an ear tag, and morphometric information was recorded (Table 1). Urine was obtained from these rats and tested for *L. interrogans* by real-time PCR. During the intervention, we euthanized trapped rats; in control sites and flanking blocks, capture-release continued, and rats were not euthanized. The University of British Columbia’s Animal Care Committee (A14-0265) approved all procedures (online Technical Appendix, https://wwwnc.cdc.gov/EID/article/24/2/17-1371-Techapp1.pdf).

We used mixed-effects multivariable logistic regression to estimate the effect of the intervention on the odds that rats carried *L. interrogans*, while controlling for clustering by city block (4). The outcome was the *L. interrogans* PCR status (negative or positive) of individual rats. The predictor variable categorized rats by block and period of capture: 0, rats caught before the intervention; 1, rats caught after the intervention in control blocks; 2, rats caught after the intervention in control blocks; and 3, rats caught after the intervention in intervention blocks. Although we did not undertake the intervention in control sites, we considered the third 2-week trapping period independently from the other trapping periods in control sites to detect any temporal changes in *L. interrogans* prevalence not associated with the intervention. We excluded the 7 rats captured both before and after the intervention to avoid double-counting individual rats. For rats recaptured within the same period as their first capture (either before or after the intervention), we averaged weight and length across captures. We also excluded 1 rat missing data for covariates (Table 1). We kept covariates, selected on the basis of their potential to confound the relationship between the intervention and *L. interrogans* carriage, in the model if they changed the estimated relationship between the predictor and outcome variables by ≥10%. Because

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length and weight were collinear, we used the covariate with the largest effect on the relationship between the predictor and outcome. We dichotomized weight around its median because it was not linear with the log-odds of the outcome. For statistical analyses, we used RStudio (Boston, MA, USA).

Of the 438 rats trapped, we included 430 in the modeling process (Table 1). Sixty-four (14.9%; 95% CI 11.7%–18.7%) rats were PCR-positive for *L. interrogans*. Of 131 rats recaptured, 5 were *L. interrogans* positive at their first capture and recapture; no positive rats changed pathogen status within a trapping period. Rats caught in intervention blocks after an intervention had 9.55 times the odds of carrying *L. interrogans* than did rats trapped before an intervention, while adjusting for weight and wound presence variables (Table 2). We found no significant changes in either flanking blocks or control blocks. In this model, 52.6% of the total model variance...
was due to the random effect of the block (11). Relooping the final model including animals that were caught both before and after the intervention did not substantially affect the results (effect of the intervention in intervention blocks; adjusted odds ratio 8.88, 95% CI 1.68–68.08).

Conclusions
This study showed that kill-trapping was associated with increased odds that rats carried *L. interrogans* in the city blocks where trapping occurred. We did not observe this effect in control blocks or nonkill flanking-blocks.

Increased intraspecific transmission of *L. interrogans* resulting from kill-trapping is a plausible explanation for the observed effect. Previous research suggests that rat-to-rat transmission of *L. interrogans* is associated with social structures in rat colonies (4). Given that culling is ineffective at removing entire rat populations (7,8,12), kill-trapping may have disrupted social structures and promoted new interactions that facilitated transmission among remaining rats. For example, culling may have removed dominant rats (13), subsequently increasing aggressive interactions among the remaining rats as they established a

Table 1. Distributions of covariates by rat-trapping period and *Leptospira interrogans* real-time PCR status, Vancouver, British Columbia, Canada, June 2016–January 2017

| Covariate                        | Total | PCR status before intervention | PCR status after intervention |
|----------------------------------|-------|---------------------------------|--------------------------------|
|                                  |       | Negative | Positive | Negative | Positive |
| Total                            | 430   | 226      | 39       | 140      | 25       |
| Season, no. (%)                  |       |          |          |          |          |
| Summer, Jun–Aug                  | 115 (27) | 83 (37) | 15 (38) | 13 (9)   | 4 (16)   |
| Fall, Sep–Nov                    | 203 (47) | 143 (63) | 24 (62) | 33 (24) | 3 (12)   |
| Winter, Dec–Feb                  | 112 (26) | 0†       | 0†       | 94 (67) | 18 (72)  |
| Sex, no. (%)                     |       |          |          |          |          |
| F                                | 205 (48) | 107 (47) | 16 (41) | 69 (49) | 13 (52)  |
| M                                | 225 (52) | 119 (53) | 23 (59) | 71 (51) | 12 (48)  |
| Sexual maturity, no. (%)         |       |          |          |          |          |
| Juvenile                         | 178 (41) | 117 (52) | 1 (3)   | 56 (40) | 4 (16)   |
| Mature                           | 252 (59) | 109 (48) | 38 (97) | 84 (60) | 21 (84)  |
| Continuous median length, cm (IQR) | 31 (26–39) | 29 (25–37) | 41 (36–43) | 30 (26–36) | 39 (33–42) |
| Wounds, no. (%)                  |       |          |          |          |          |
| No                               | 316 (73) | 173 (77) | 11 (28) | 115 (82) | 17 (68) |
| Yes                              | 114 (27) | 53 (23) | 28 (72) | 25 (18) | 8 (32)   |
| Weight, g, no. (%)               |       |          |          |          |          |
| <122                             | 212 (49) | 129 (57) | 1 (3) | 77 (55) | 5 (20) |
| ≥122                             | 218 (51) | 97 (43) | 38 (97) | 63 (45) | 20 (80) |
| *IQR, interquartile range.*      |       |          |          |          |          |
| †No periods before the intervention period were conducted during winter.

Table 2. Results of model building in a study of the effects of culling on *Leptospira interrogans* carriage by rats, Vancouver, British Columbia, Canada, June 2016–January 2017

| Covariate                        | Unadjusted odds ratio* (95% CI) | Adjusted odds ration† (95% CI) | p value |
|----------------------------------|----------------------------------|--------------------------------|---------|
| Season                           | Reference                        | –                               | –       |
| Summer                           | 0.44 (0.13–1.39)                | –                               | –       |
| Fall                             | 0.87 (0.22–3.24)                | –                               | –       |
| Winter                           | 1.28 (0.70–2.37)                | –                               | –       |
| Sex                              | Reference                        | –                               | –       |
| F                                | 1.81 (1.42–2.39)                | 3.67 (1.73–9.12)                | 0.0013  |
| M                                | 16.26 (6.28–51.95)              | –                               | –       |
| Sexual maturity                  | Reference                        | –                               | –       |
| Juvenile                         | 1.25 (1.18–1.35)                | –                               | –       |
| Mature                           | 17.88 (7.22–53.28)              | 9.98 (3.70–31.74)               | <10⁻⁴   |
| Continuous length, cm            | Reference                        | Reference                       | Reference |
| Weight, g                        | Reference                        | Reference                       | Reference |
| <122                             | 0.69 (0.22–2.00)                | 0.77 (0.22–2.58)                | 0.68    |
| ≥122                             | 1.50 (0.49–4.40)                | 2.22 (0.65–7.47)                | 0.19    |
| Intervention                     | Reference                        | Reference                       | Reference |
| Before intervention, all block types, n = 261 | – | – | – |
| After intervention, control blocks, n = 97 | 0.69 (0.22–2.00) | 0.77 (0.22–2.58) | 0.68 |
| After intervention, nonkill flanking blocks, n = 33 | 1.50 (0.49–4.40) | 2.22 (0.65–7.47) | 0.19 |
| After intervention, intervention blocks, n = 39 | 8.67 (2.02–55.00) | 9.55 (1.75–78.31) | 0.016 |

*Univariable relationships between the indicated covariate and *L. interrogans* status, while controlling for the random effect of the block.
†Results of the final multivariable model in which the effect of each covariate is adjusted for other covariates in the model.
‡Dashes indicate variables not carried forward into the final multivariable model on the basis of statistical confounding criteria.
new social hierarchy. The positive association between L. interrogans status and weight/wound presence (which are correlated with hierarchical dominance) supports this hypothesis because the bacteria may be transmitted through specific aggressive/dominance interactions (4).

We assessed only the effect of culling on a single rat-borne pathogen. L. interrogans might be particularly susceptible to the effects of culling because of its dependence on rat social structures. Other vectorborne (e.g., flea-borne Rickettsia spp. [14]) or environmentally acquired (e.g., methicillin-resistant Staphylococcus aureus [15]) rat-associated pathogens might not be as easily influenced by culling. Future studies should determine the duration of effects induced by lethal control because effects on L. interrogans prevalence may wane with time. However, given that such methods are ineffective at removing entire rat populations and might therefore be used repeatedly as the population rebounds (7,8,12), short-term effects may be particularly important.

We demonstrated that rat culling has the potential to increase the odds for L. interrogans carriage among remaining rats and thus could potentially increase the risk for transmission to humans. Although public health risks resulting from such an increase postintervention might be offset by a decrease in the number of rats, we were unable to quantify the size of the rat population before and after intervention. Practical and ethical considerations make it difficult to empirically demonstrate a direct link between culling and increased pathogen transmission from rats to humans. Rather, after culling, the potential for a person to encounter a rat carrying L. interrogans increases if a person encounters a rat, suggesting that the risk for zoonotic transmission increases per rat contact.

The convergence of this study with previous literature documenting that reactive culling is often unsuccessful at removing rat populations (7,8,12) indicates that such methods are ineffective. Instead, ecologically based rodent management, which focuses on reducing resources available to rats (8), should be more widely applied to urban environments.

By integrating our results with other studies on the impacts of culling wild animals to control communicable diseases (5,6), we can conclude that killing animal reservoirs of human pathogens might have unintended consequences on the disease risks. This hypothesis underscores the importance of understanding the ecology of the targeted animal reservoir to design effective control programs.

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Technical Appendix

Real-Time PCR

Nucleic acid from rat urine was extracted with a 96-well magnetic particle processor using the MagMAX Pathogen RNA/DNA kit (Life Technologies). Extractions were performed following the procedure outlined by the manufacturer for low-cell content samples. Starting urine volume ranges were 20–200 µL, and all were volume corrected to 200 µL using sterile, 1× phosphate buffer solution buffer, pH 7.4.

Nucleic acid extracts were amplified using a real-time PCR (Life Technologies) that targets the LipL32 gene (encodes an outer membrane lipoprotein virulence factor [1] of pathogenic *Leptospira* spp.). Real-time PCR was performed using the Agpath-ID One-Step real-time PCR Kit (Life Technologies). A Taqman exogenous internal positive control (IPC) (Life Technologies) was also run to ensure that there was no PCR inhibition due to the inhibitory nature of urine samples.

Each 25-µL reaction contained 2× real-time PCR buffer, 25X real-time PCR enzyme, 800 nM each of forward primer (5′-AAG CAT TAC CGC TTG TGG TG-3′) and reverse primer (5′-GAA CTC CCA TTT CAG CGA TT-3′), 200-nM probe (5′-FAM/AAA GCC AGG ACA AGC GCC G/BHQ1-3′), 10X Exo IPC Mix, 500× Exo IPC DNA (diluted 10-fold), nuclease-free water and 5 µL of DNA template. The reaction was incubated at 50°C for 2 min, 95°C for 10 min, and then amplified for 45 cycles at 95°C for 15 s, 58°C for 1 min. Samples were run on an ABI7500 Fast PCR system (Life Technologies) and analyzed using the SDS software version 1.4 (Life Technologies).

*Leptospira* spp. primers and probe were made by Integrated DNA Technologies (San Diego, CA, USA). A negative extraction control, negative template control, and 2 positive
amplification controls were used per real-time PCR run. The positive control was *L. interrogans*, serovar *copenhageni* (Tim Witchell, University of Victoria, BC, Canada, June 2012).

**Field Methods**

**Trapping**

Trapping was conducted during June 2016–January 2017. Ten Tomahawk Rigid Traps (Tomahawk Live Traps, Hazelhurst, WI, USA) were placed in the alley that bisected each city block. To prevent vandalism, traps were fitted into stainless steel trap covers (Integrated Pest Supplies Ltd, New Westminster, BC, Canada) and chained to immovable objects. Traps were baited with peanut butter mixed with oats. Hydrogel (ClearH2O, Westbrook, ME, USA) was provided as a water source.

Three study sites were trapped at a time, such that 90 traps were deployed in 9 city blocks at any given point during this study. Prebaiting, in which cages were fixed open and baited, was conducted for 1 week before any new trapping period to acclimatize rats to cages. During trapping periods, traps were set each evening by 4 PM and checked each morning by 7 AM, 5 days a week. On the sixth and seventh days, traps were fixed open and baited. Traps and associated equipment were sanitized in 10% bleach and/or 70% ethanol (2) after coming into contact with any rat and after any period of prebaiting.

**Sample Collection**

Captured rats were transported to the back of a mobile laboratory van and given Hydrogel to promote urination, and their cages were covered with a blanket to minimize stress until sampling. Urine was obtained by placing caged rats directly above a bleach-sanitized plastic tray until they urinated into it. Urine was collected using a sterile syringe and was stored at ~80°C until analysis. Subsequently, rats were transferred into an inhalation induction chamber (Kent Scientific, Torrington, CT, USA) and anesthetized with 5% isoflurane in oxygen using an isoflurane vaporizer (Associated Respiratory Veterinary Services, Lacombe, AB, Canada). Anesthesia was maintained throughout sampling.

Each rat was given a unique laser-etched ear-tag (Kent Scientific) for identification upon recapture. The following demographic and morphometric characteristics were assessed: body weight (grams), total length (nose-to-tail in centimeters), sexual maturity (males with scrotal
testes and females with a perforate vagina were considered mature), sex (male or female), and the presence/absence of bite wounds (presence determined in accordance with [3]). Rats were allowed to recover fully from anesthesia before being released at the exact location of their capture ≈15–30 minutes after sampling.

Rats that had been previously captured and sampled were resampled if >7 days had passed since their previous capture. One week was determined to be an appropriate interval in which to detect a change in *L. interrogans* infection status because the bacterium can be detected in renal tissue and rat urine in as little as 1 week after experimental infection (4,5). Rats caught in intervention blocks during the 2-week kill-trapping period were anesthetized using isoflurane and euthanized by intracardiac injection with pentobarbital.

**Interactions**

We explored biologically plausible interactions between the effect of the intervention and covariates by running the final multivariable model in strata of each covariate. For example, we hypothesized that effect of the intervention on *L. interrogans* carriage might be modified by socially relevant morphologic characteristics, such as sexual maturity, given that transmission between rats might depend on social structures (6). For sexual maturity, we therefore stratified and recomputed the multivariable model for juvenile and mature rats separately. However, because of a limited sample size, there was not enough statistical power to test for interactions in the restricted strata.

Of all covariates tested, sex was the only variable that had large enough strata to test for interactions. These data indicated that the effect of the intervention may be more pronounced among females (adjusted odds ratio 7.62, 95% CI 0.81–110.49) than among males (adjusted odds ratio 3.71, 95% CI 0.38–63.45), while weight and bite wounds remained constant. However, the effect of the intervention was not significant in either strata, again, most likely because of the limited sample size in each group.

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