The impact of comorbidity on mortality in multiple myeloma
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The survival of multiple myeloma patients has improved over the last two decades with the introduction of new drugs [1, 2]. In addition, the concept of combined therapies has evolved in pursuit of synergistic effects [3]. Although new treatment options have led to improved survival, multiple myeloma remains an incurable malignant disease and an important challenge is to apply data from clinical trials to the real-life population of multiple myeloma [4].

The incidence of multiple myeloma increases with age and the prolonged survival of the population in general will lead to an increased number of multiple myeloma
patients in the future [5]. Comorbidities correlate with aging and make the elderly particularly vulnerable to toxicities of therapy. A key issue is therefore to choose the optimal therapy for these patients [6]. An increasing awareness of the vulnerability of the elderly has lead to recommendations for dose adjustment of treatment in patients with comorbid conditions [7]. Recently, a frailty score was proposed by the International Myeloma Working Group including age, geriatric assessment, and comorbidity assessed by the Charlson Comorbidity Index [8]. This scoring system was based on data from three prospective clinical trials and predicts mortality and the risk of toxicity in myeloma patients. However, patients with comorbidities are often excluded from clinical trials and therefore these data may not necessarily provide guidance for treatment decisions for the elderly patients in general. Data on the prevalence and impact of comorbidity in the real-life population of multiple myeloma patients is limited [5, 9].

We therefore conducted a population-based Danish study of all patients with newly diagnosed multiple myeloma and a matched control population to compare the burden of comorbidity and to evaluate the impact of comorbidity on mortality.

Materials and Methods

Denmark has approximately 5.6 million inhabitants and every resident is registered in the national Civil Registration System (CRS), and is assigned a personal identification number (the CPR-number) at birth, which allows linkage between demographic and medical registries [10]. The National Health Service provides tax-supported health care for all citizens in Denmark and guarantees free access to hospitals.

Study population

The Danish Multiple Myeloma Registry was established in 2005 by the Danish Myeloma Study Group (DMSG) [11]. Danish hematology departments are obligated to report all incident cases of multiple myeloma, solitary plasmacytoma, MGUS, and plasma cell leukemia to the registry. The database include the date of myeloma diagnosis, percentage of plasma cells in the bone marrow, result of skeletal X-ray, ECOG performance status, clinical presentation, for example, spinal cord compression, type and concentration of the M-component, and basic biochemical parameters, for example, beta-2-microglobulin. In the study period, the 2003 criteria for classification of monoclonal gammopathies from the International Myeloma Working Group were used [12]. Annual links to the Danish National Patient Registry ensure that cases of multiple myeloma that have not been reported to the Danish Multiple Myeloma Registry are subsequently included. A recent validation of the registry has shown almost 100% completeness [11].

Data on comorbidity

The Danish National Patient Registry (DNPR) has, since 1977, routinely collected nationwide data on all nonpsychiatric hospitalizations, and since 1995, visits to emergency departments and outpatient clinics have also been included [13]. The DNPR records CPR-number and date of each hospital visit, together with primary and secondary discharge diagnoses. The DNPR has almost complete coverage of national hospital admissions [14]. Data have been coded according to the International Classification of Diseases, 10th revision (ICD-10) since 1994.

Comorbidity was classified according to the 19 diseases included in the Charlson Comorbidity Index (CCI) [15]. We computed a CCI score based on registered diagnoses within 10 years and until 1 month before diagnosis of multiple myeloma. Four levels of comorbidity were defined: 0 (“low”) for individuals with no recorded underlying diseases included in the CCI; 1 (“moderate low”), 2 (“moderate high”), and 3 (“high”).

General population comparison cohort

For each myeloma patient, we randomly chose 10 members of the general population from the CRS, matched by year of birth and sex. The members of the comparison cohort had to be alive with no diagnosis of multiple myeloma on the date when follow-up of the matched myeloma patient began. Thus, the control persons were included at the time of diagnosis of their corresponding multiple myeloma patient with whom they were matched. Data on comorbidity was obtained from the DNPR.

Statistical analysis

Comparisons of binary variables were conducted by chi-square test or Fisher’s exact test, mid-P approach [16]. Continuous variables were presented using medians with range or interquartile range. Ordinal variables were compared using Mann–Whitney test. Overall survival was defined as the time from diagnosis to death from any cause and was described using the Kaplan–Meier method. Comorbidity was assessed for two different time periods, namely from 1 year until 1 month before diagnosis of multiple myeloma and from 10 years until 1 month before diagnosis of multiple myeloma. Cox proportional hazards regression models were performed to evaluate the prognostic significance of comorbidity and the results are
presented as estimated hazard ratios (HRs) with 95% confidence intervals. The proportionality assumptions for the models were assessed by scaled Schoenfeld residuals and we did not find violations of the assumptions. When comparing subgroups with few events, a permutation method was used for calculating \( P \)-values [17]. R function “permlogrank” in the “clinfun” package was used for these calculations. The interaction between myeloma and comorbidities on mortality was inspected by calculating an interaction contrast on an additive scale. The age/gender standardized mortality rate differences between multiple myeloma and population controls for each level of comorbidity were calculated and compared to the difference in multiple myeloma/population controls without comorbidities [18]. For calculations of nonlinear effects on survival of age, a general additive model was performed with age as a smoothing spline with four degrees of freedom. All tests were two-sided and \( P \)-values of 0.05 were considered statistically significant. All confidence intervals are 95% and two-sided. Data analyses were performed using R version 3.2.3 (R Foundation for Statistical Computing, Vienna, Austria).

The study was approved by the Danish Data Protection Agency (No. 2008-58-0028).

Results

The study included a total of 2190 patients with symptomatic multiple myeloma in the period 2005–2012 in Denmark. The median age at diagnosis was 70 years (range 30–98 years), and 1204 (55%) were men. The median duration of follow-up for patients alive was 4.3 years (interquartile range 2.4–6.3). Characteristics of the patients and population controls are described in Table 1.

Comorbidity in the myeloma cohort and population controls

Comorbidity at cohort entry was registered in 40.9% of the patients. The overall comorbidity was increased in multiple myeloma patients compared with population controls, odds ratio (OR) 1.4 (1.1–1.7). The prevalence of comorbidity was higher in patients over 65 years of age compared with the younger patients, for example,
three or more comorbidities in 226 (15.9%) patients >65 years versus 47 (6.1%) patients ≤65 years (P < 0.0001). In addition, comorbidity was higher in male patients (42.3%) compared with female patients (39.3%) (P = 0.04). Table 2 describes details on registration of the diseases in the Charlson Comorbidity Index. The registration of comorbidity was highly increased within the year preceding diagnosis of multiple myeloma (OR 3.0 [2.5–3.5; P < 0.0001]) which was attributable to an increased registration of various diseases, in particular, renal disease with OR 11.0 (8.1–14.9) (Table 2). By contrast, comorbidity was not increased in the time period from 10 years until 1 year before diagnosis of multiple myeloma (OR 1.0 [0.9–1.1, P = 0.77]) and the only deviations of individual diseases in this time period were an increased prevalence of renal disease (OR 1.5 [1.1–2.1; P = 0.02]) and lower prevalence of dementia (OR 0.3 [0.2–0.6; P = 0.0001]). An increased registration of venous thromboembolism was observed in the period from 10 years until diagnosis of multiple myeloma and in the year preceding the diagnosis of multiple myeloma with OR of 1.5 (1.1–1.9; P = 0.003) and 2.8 (1.6–4.5; P = 0.0004), respectively.

**Survival in the multiple myeloma cohort**

The median overall survival in all patients was 3.3 years (1.0–6.5) comprising median survival of 5.7 years (2.5 – not reached) and 2.3 years (0.6–4.7) in patients aged 65 years or younger and in patients older than 65 years, respectively. Figure 1 shows the survival stratified according to the Charlson Comorbidity Index for

### Table 2. Hazard ratios comparing the individual diseases included in the Charlson Comorbidity Index in 2190 Danish patients with newly diagnosed symptomatic multiple myeloma and 21,900 population controls.

| Diagnosis                                      | Registered comorbidity from 10 years until 1 month before diagnosis of multiple myeloma | Registered comorbidity from 1 year until 1 month before diagnosis of multiple myeloma |
|------------------------------------------------|----------------------------------------------------------------------------------------|--------------------------------------------------------------------------------------|
| Any Charlson condition                        | 896 (40.9%) 1.4 (1.1–1.7) <0.0001                                                    | 194 (8.9%) 3.0 (2.5–3.5) <0.0001                                                    |
| Myocardial infarction                         | 118 (5.4%) 1.1 (0.9–1.3) 0.5745                                                       | 18 (0.8%) 1.7 (1.0–2.7) 0.0557                                                      |
| Congestive heart failure                      | 126 (5.8%) 1.4 (1.1–1.7) 0.0016                                                       | 36 (1.6%) 2.8 (1.9–4.0) <0.0001                                                    |
| Peripheral vascular disease                   | 82 (3.7%) 0.9 (0.7–1.1) 0.3126                                                       | 16 (0.7%) 1.2 (0.7–1.9) 0.5429                                                     |
| Cerebrovascular disease                       | 160 (7.3%) 0.9 (0.7–1.0) 0.1390                                                       | 29 (1.3%) 1.3 (0.9–1.9) 0.2181                                                     |
| Dementia                                       | 17 (0.8%) 0.5 (0.3–0.8) 0.0015                                                       | 8 (0.4%) 1.1 (0.5–2.1) 0.8539                                                      |
| Chronic pulmonary disease                     | 147 (6.7%) 1.0 (0.8–1.2) 0.9526                                                       | 28 (1.3%) 1.7 (1.1–2.5) 0.0140                                                    |
| Connective tissue disease                     | 72 (3.3%) 1.2 (0.9–1.5) 0.2264                                                       | 19 (0.9%) 3.9 (2.2–6.6) <0.0001                                                    |
| Ulcer disease                                 | 89 (4.1%) 1.6 (1.3–2.0) 0.0002                                                       | 33 (1.5%) 5.3 (3.4–8.1) <0.0001                                                    |
| Mild liver disease                            | 17 (0.8%) 1.2 (0.7–2.0) 0.4267                                                       | 6 (0.3%) 3.2 (1.2–7.7) 0.0276                                                     |
| Diabetes Mellitus                             | 71 (3.2%) 1.0 (0.8–1.3) 0.7061                                                       | 12 (0.5%) 1.6 (0.8–2.8) 0.1664                                                    |
| Hémiplegia                                     | 7 (0.3%) 1.3 (0.6–2.8) 0.4779                                                       | 1 (0.0%) 2.2 (0.1–14.5) 0.5324                                                    |
| Moderate and severe renal disease             | 131 (6.0%) 3.7 (3.0–4.6) <0.0001                                                    | 89 (4.1%) 11.0 (8.1–14.9) <0.0001                                                 |
| Diabetes mellitus with chronic complications  | 73 (3.3%) 1.1 (0.9–1.4) 0.3096                                                       | 18 (0.8%) 2.3 (1.3–3.7) 0.0042                                                    |
| Any tumor                                      | 223 (10.2%) 1.2 (1.0–1.4) 0.0149                                                      | 49 (2.2%) 1.9 (1.3–2.5) 0.0003                                                    |
| Leukemia                                       | 8 (0.4%) 1.1 (0.5–2.1) 0.8539                                                       | 2 (0.1%) 1.6 (0.2–6.0) 0.5517                                                     |
| Lymphoma                                       | 21 (1.0%) 2.0 (1.2–3.2) 0.0069                                                       | 8 (0.4%) 5.8 (2.3–13.6) 0.001                                                    |
| Moderate and severe liver disease             | 2 (0.1%) 0.5 (0.1–1.7) 0.3254                                                       | 1 (0.0%) 1.9 (0.1–11.3) 0.6145                                                    |
| Metastatic solid tumor                        | 41 (1.9%) 1.9 (1.3–2.6) 0.0006                                                       | 25 (1.1%) 4.1 (2.6–6.5) <0.0001                                                   |
| AIDS                                           | 0 (0.0%) – 0.62 0.0006                                                              | 0 (0.0%) – – 0.0006                                                              |
myeloma patients and population controls. Patients with any registered comorbidity had increased mortality compared with patients without comorbidity, HR 1.6 (1.5–1.8). The mortality according to the diseases that constitute the Charlson Comorbidity Index is shown in Table 3. Registration of venous thromboembolism was associated with increased mortality (HR 2.0 [1.5–2.7; \( P < 0.0001 \))).

Risk factors for survival

A Cox proportional hazards model was used to analyze the impact on mortality of comorbidity and well-known risk factors, including age, International Staging System, serum lactate dehydrogenase, serum creatinine, serum C-reactive protein, IgA M-component, and WHO performance status (0–1 vs. 2–4) (Table 4). The mortality was increased in patients registered with comorbidity according to Charlson Comorbidity Index compared with patients without registered comorbidity, for example, hazard ratios of 1 (ref), 1.5 (1.2–1.8), 1.3 (1.0–1.5), and 1.7 (1.4–2.1) in the comorbidity groups 0, 1, 2, and ≥3, respectively. A division of comorbidity into two groups similar to the International Myeloma Working Group (IMWG) frailty study (CCI <2 vs. CCI ≥2) showed an increased mortality in the group with high comorbidity (HR 1.3 [1.1–1.5; \( P = 0.0003 \)]) [8].

Comorbidity and WHO performance status

There was no association between Charlson Comorbidity Index and the WHO performance status at time of diagnosis of multiple myeloma (\( p = 0.07 \), chi-square test) (Table 5).

High-dose melphalan with hematopoietic stem cell support

High-dose melphalan with haematopoietic stem cell support (HDT) was used as first-line therapy in 582 (76.1%) patients 65 years or younger, and in 103 (7.2%) patients older than 65 years. Comorbidity affected the utilization of HDT, for example, used in 452 (81.1%) patients 65 years or younger with no registered comorbidity in contrast to 24 (51.1%) patients with three or more comorbidities (Table 6).
In this large nationwide cohort study, we found an increased comorbidity in patients with multiple myeloma at time of diagnosis compared to population controls. This increase in comorbidity was mainly confined to the year preceding diagnosis of myeloma. In both younger and elderly
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Comorbidity in multiple myeloma patients, comorbidity was associated with increased mortality.

The prevalence of comorbid diseases in our study is in accordance with the observations in a Swedish population-based study based on the Swedish Cancer Registry where comorbidity was seen in approximately 40% of the multiple myeloma patients and with an Italian study based on a regional multiple myeloma registry [9, 19]. We provide a further extension of the importance of comorbidity by comparing comorbidity in multiple myeloma patients with population controls and we found an increased registration of various comorbidities in multiple myeloma patients, in particular renal disease. This finding is consistent with studies on solid cancers that observed an increased comorbidity compared to the background population at diagnosis in specific cancers, for example, in lung and colorectal cancer patients [8, 20]. In our study, the increased comorbidity was mainly caused by an increased registration of various comorbidities in multiple myeloma patients, in particular renal disease. This finding is consistent with studies on solid cancers that observed an increased comorbidity compared to the background population at diagnosis in specific cancers, for example, in lung and colorectal cancer patients [8, 20]. In our study, the increased comorbidity was mainly caused by an increased registration of various comorbidities within the last year preceding the diagnosis of multiple myeloma. Several factors may be important and involved in this observation. Well-known multiple myeloma complications and morbidity secondary to multiple myeloma may partly explain the increased frequency, for example, the increased risk of myocardial infarction within the first years might reflect the increased risk of thromboembolism observed in multiple myeloma [21]. This is supported by the fact that we found the same pattern for venous thromboembolism in our study cohort. The high prevalence of renal disease compared to the population controls is plausible and reflects that renal impairment is a frequent complication in multiple myeloma [22]. In addition, our cohort will have included cases of monoclonal gammopathy of renal significance (MGRS) that later underwent malignant transformation to multiple myeloma [23, 24]. An essential issue is the distinction between multiple myeloma complications and true comorbidity. A considerable part of the registered comorbidity within the last year prior to diagnosis might be related to the diagnostic process rather than reflecting true comorbidity in multiple myeloma. In a few cases, it might reflect the detection of multiple myeloma in patients in diagnostic workup for unrelated diseases which are subsequently registered as part of the comorbidity. However, symptomatic multiple myeloma is not likely to pass undetected for longer periods of time due to the prominent symptomatology. The association between registration of comorbidity and time to diagnosis of multiple myeloma raises methodological issues regarding when to assess comorbidity in studies on multiple myeloma.

We found that comorbidity increased the mortality in multiple myeloma patients. There was no clear difference in mortality between the subgroups 1, 2, and ≥3 according to Charlson Comorbidity Index. However, differences in mortality were observed when comorbidity according to Charlson Comorbidity Index was used as

### Table 5. Association between comorbidity and WHO performance status in 2190 Danish patients with newly diagnosed symptomatic multiple myeloma.

| Charlson comorbidity Index | WHO performance status 0–1 | WHO performance status 2–4 |
|---------------------------|-----------------------------|-----------------------------|
| 0                         | 838 (65.0%)                 | 451 (35.0%)                 |
| 1                         | 191 (60.6%)                 | 124 (39.4%)                 |
| 2                         | 204 (66.9%)                 | 101 (33.1%)                 |
| ≥3                        | 157 (58.1%)                 | 113 (41.9%)                 |

*P* = 0.07 (chi-square test). Numbers do not add up to 2190 due to missing data on WHO performance status.

### Table 6. Use of treatment with high-dose melphalan with stem cell support (HDT) according to comorbidity in 2190 Danish patients with newly diagnosed symptomatic multiple myeloma.

| Comorbidity | All patients | Patients age ≤65 years |
|-------------|--------------|------------------------|
|             | HDT          | Other treatment        | HDT          | Other treatment |
| 0           | 521 (76.1%)  | 773 (51.4%)            | 452 (77.7%)  | 105 (57.4%)     |
| 1           | 64 (9.3%)    | 253 (16.8%)            | 49 (8.4%)    | 28 (15.3%)      |
| 2           | 73 (10.7%)   | 233 (15.9%)            | 58 (10.0%)   | 26 (14.2%)      |
| ≥3          | 27 (3.9%)    | 246 (16.3%)            | 23 (4.0%)    | 24 (13.1)       |

*Comorbidity according to Charlson Comorbidity index. Chi-square test. HDT: intention-to-treat. HDT versus no HDT for all patients and for patients with age ≤65 years.*
dichotomous variables, for example, between patients with no registered comorbidity and patients with any grade of comorbidity, or between patients with comorbidity score less than 2 compared with patients with a score of 2 or higher. The HR of 1.3 in this later division is in accordance with the HR of 1.37 found in the IMWG frailty study by Palumbo et al. [8]. This frailty score is based on data from clinical trials and due to study exclusion criteria, the pattern of comorbidity is likely to be different from that of unselected myeloma patients in a population-based study. Thus, our study provides support for the comorbidity element of the IMWG frailty score also being valid in a population-based setting. However, there is a need for studies that validate the frailty score and address comorbidity in population-based settings.

Charlson Comorbidity Index creates a pooled estimate of comorbidity and assigns particular weights to a number of diseases [15]. It was developed almost 30 years ago to predict mortality in medical patients admitted to hospitals and has been shown to be robust in a number of clinical settings, including validated for several different cancers and it is the most extensively studied comorbidity index for predicting mortality. However, the prognosis of the individual included diseases has changed over time, for example, the prognosis of AIDS has considerably changed since the eighties and some of the comorbidities are rarely detected in myeloma patient, for example, dementia and AIDS. In addition, the Charlson Comorbidity Index only includes limited information on disease severity. A recent study indicates that the Freiburg comorbidity index which is based on performance status, renal impairment, and lung disease might be more useful than the Charlson Comorbidity Index in predicting overall survival in elderly patients with multiple myeloma [25].

The adverse prognostic impact of comorbidity on survival is complex. Comorbidities may independently increase the risk of death or might add to well-known myeloma risk factors, for example, further increase the risk of serious infections or venous thromboembolic disease. In addition, comorbidity may indirectly affect the prognosis by affecting the choice of myeloma treatment and lead to more frequent dose reduction [26]. Our data confirm how a high burden of comorbidity reduces the use of HDT in younger myeloma patients. Our study design did not permit an evaluation of the association between comorbidity and the use of other myeloma treatment and dose reductions.

The main strength of our study is its large size with inclusion of all newly diagnosed multiple myeloma patients in a population-based setting and the use of a population comparison cohort. The risk of selection bias in this setting is negligible which contrasts studies including patients participating in clinical trials or patient series from single centers. We included the very old and frail patients and our study reflects the real-life population of multiple myeloma patients in Denmark. Consequently, the mortality of the elderly patients is higher than reported in most other studies [27–29]. Furthermore, the use of DNPR data ensures a complete nationwide ascertainment of former comorbid conditions in the myeloma patients and the comparison cohort. Despite these advantages, our registry-based study has some limitations. We had data on ISS stage in most patients, whereas cytogenetic and FISH only were performed in a limited number of the patients, in particular, in the early part of the inclusion period. Consequently, we were not able to include this important prognostic factor in our analysis of factors affecting mortality. The DNPR provides limited data on the severity and duration of the comorbid conditions and we were not able to validate the proposed Freiburg index. However, the positive predictive values in the DNPR of the diagnosis groups included in the Charlson Comorbidity Index are high (overall 98.0%) [30]. Diseases and complications diagnosed in primary health care are not registered in the DNPR and this under-registration could diminish potential effects of specific conditions. Another concern is the risk of coding errors leading to misclassification of comorbidity. However, this bias is likely to be similar in the myeloma patients and populations controls at the time of interest and not subject to the surveillance bias that might follow the diagnosis of multiple myeloma.

In conclusion, we found increased registration of comorbid disease within the year prior to diagnosis of multiple myeloma. This finding is likely to represent complications and morbidity secondary to yet undiagnosed multiple myeloma and also conditions related to the diagnostic process but it underlines the potential burden of comorbidity in myeloma. We found that comorbidity increased the mortality in multiple myeloma patients when we assessed the comorbidity in the Charlson Comorbidity Index as dichotomous variables, for example, no comorbidity versus any comorbidity or comorbidity score less than 2 compared with a score of 2 or higher which supported the use of the index in the IMWG frailty study. However, we did not find any clear difference in mortality between the subgroups 1, 2, and ≥3 according to Charlson Comorbidity Index.

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Conflict of Interest

None declared.

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