Breast Cancer Survival Analysis: Applying the Generalized Gamma Distribution under Different Conditions of the Proportional Hazards and Accelerated Failure Time Assumptions

Alireza Abadi, Farzaneh Amanpour¹, Chris Bajdik², Parvin Yavari³

ABSTRACT

Background: The goal of this study is to extend the applications of parametric survival models so that they include cases in which accelerated failure time (AFT) assumption is not satisfied, and examine parametric and semiparametric models under different proportional hazards (PH) and AFT assumptions.

Methods: The data for 12,531 women diagnosed with breast cancer in British Columbia, Canada, during 1990–1999 were divided into eight groups according to patients’ ages and stage of disease, and each group was assumed to have different AFT and PH assumptions. For parametric models, we fitted the saturated generalized gamma (GG) distribution, and compared this with the conventional AFT model. Using a likelihood ratio statistic, both models were compared to the simpler forms including the Weibull and lognormal. For semiparametric models, either Cox’s PH model or stratified Cox model was fitted according to the PH assumption and tested using Schoenfeld residuals. The GG family was compared to the log-logistic model using Akaike information criterion (AIC) and Baysian information criterion (BIC).

Results: When PH and AFT assumptions were satisfied, semiparametric and parametric models both provided valid descriptions of breast cancer patient survival. When PH assumption was not satisfied but AFT condition held, the parametric models performed better than the stratified Cox model. When neither the PH nor the AFT assumptions were met, the log normal distribution provided a reasonable fit.

Conclusions: When both the PH and AFT assumptions are satisfied, the parametric and semiparametric models provide complementary information. When PH assumption is not satisfied, the parametric models should be considered, whether the AFT assumption is met or not.

Key words: Breast cancer, generalized gamma distribution, parametric regression, stratified Cox model, survival analysis

INTRODUCTION

The Cox proportional hazards (PH) model is popular for analyzing survival data. The utility of this model stems from the
fact that few assumptions are needed to determine hazard ratios based on the coefficients. The coefficients are easily interpreted and clinically meaningful.[1] The “stratified Cox (SC) model” is a modification of the Cox PH model, which allows for control by stratification of a predictor that does not satisfy the PH assumption.[2]

Parametric models are used only occasionally in analyzing clinical studies of survival despite offering some advantages over semiparametric models. Parametric regression analysis is an attractive alternative to the widely used Cox model when hazard functions themselves are of primary interest, or when relative survival times are the primary measure of association.[3] When empirical information is available, parametric models can provide insight into the shape of the baseline hazard and baseline survival. However, fully parametric models involve stronger assumptions than semiparametric options. Furthermore, difficulty choosing the appropriate family of distributions leads many researchers to prefer the Cox model.[4] Some parametric models are accelerated failure time (AFT) models which assume that the relationship between the logarithm of survival time and covariates is linear.[5] Violation of the AFT assumption makes the parametric models more complicated.

One approach to address these difficulties is fitting the generalized gamma (GG) distribution. This extensive family contains nearly all commonly used distributions including the exponential ($\lambda = \sigma = 1$), Weibull ($\lambda = 1$), and log normal ($\lambda = 0$). If a general parametric distribution includes other distributions as special cases, it is a nesting (larger) family of the specific distributions.[5] The GG distribution includes three specific distributions, and thus represents a nesting family of them, allowing us to evaluate the appropriateness of the specific distributions relative to each other. Testing the appropriateness of a family of distributions is equivalent to testing whether a subset of parameters in its nesting distribution are equal to specific values, and can be performed using a likelihood ratio test.[5]

The log-logistic distribution is a commonly used distribution in survival analysis, which is not nested in the GG family. To compare the selected parameter distribution with the log-logistic distribution, we used simple procedures based on Bayesian information criterion (BIC) ($r = L(\hat{b}) - \frac{n}{2}\log(n)$); Schwarz, 1978 and Akaike information criterion (AIC; Akaike, 1969), in which $r$ is defined as $r = L(\hat{b}) - 2p$. In our comparisons, the candidate distribution with the largest $r$ value was considered the best fit.[5]

### Statistical analysis

#### Parametric model

For choosing the appropriate parametric model, we started by fitting the saturated GG distribution. This distribution is a three-parameter family with location (\(\beta\)), scale (\(\sigma > 0\)), and shape (\(\lambda\)) parameters, in which all the three parameters depend on covariates. It should be noted that the conventional AFT model holds only when covariate effects are modeled through the beta parameter. If we extend the analysis to covariates having effects through the sigma and/or lambda parameters, these are no longer conventional AFT models.[3]

The GG family contains nearly all of the most commonly used distributions in survival analysis, including the exponential (\(\lambda = \sigma = 1\)), Weibull (\(\lambda = 1\)), and log normal (\(\lambda = 0\)). If a general parametric distribution includes other distributions as special cases, the general distribution is called a nesting (larger) family of the specific distributions.[5] The GG distribution includes three specific distributions, and thus represents a nesting family of them, allowing us to evaluate the appropriateness of the specific distributions relative to each other. Testing the appropriateness of a family of distributions is equivalent to testing whether a subset of parameters in its nesting distribution are equal to specific values, and can be performed using a likelihood ratio test.[5]
For AFT models, we can estimate relative survival time by exponentiating the coefficient of a variable. For other models, we can calculate the relative time, RT (p), using appropriate formulae. The relative times are defined for 0 < p < 1 as the ratio of the corresponding quantile functions, RT (p) = t_{1}(p)/t_{0}(p). The interpretation of RT (p) is that the time required for proportion p of individuals in the exposed or treated population to experience the event of interest is RT(p)-fold the time for the same proportion of events to occur in the reference population. Links between quantiles of the gamma and GG facilitats use of software to obtain the percentiles of the GG.[3]

**Semiparametric model**

We can also describe the distribution of survival time by specifying the hazard function. The advantage of this approach is that we directly address the aging process. Cox was the first to propose the model, specifying the hazard function as a function of time and the covariates:[1]

\[ H(t, x, b) = h_0(t) \exp(xb). \]

With this parameterization, the hazard ratio is:

\[ \text{HR} (t, x_1, x_0) = \exp(b(x_1 - x_0)). \]

The Cox PH model assumes that the hazard ratio for any two specifications of predictors is constant over time, and Schoenfeld residuals can be used to assess the PH assumption.[2] The “SC model” is a modification of the Cox PH model, which allows for control by “stratification” of a predictor that does not satisfy the PH assumption. The predictor that does not satisfy the PH assumption is being adjusted by stratification, whereas the predictor that satisfies the PH assumption is being adjusted by its inclusion in the model. The hazard ratio value for the effect of variables in each stratum can be estimated. Nevertheless, the hazard ratio value for the effect of a stratified variable cannot be estimated. Furthermore, we applied the likelihood ratio (LR) test to check the interaction between stratified variable and variables in each stratum.[2]

A standard treatment protocol for breast cancer is determined mainly by the patient’s age, stage of cancer, and tumor sensitivity to certain hormones. Breast cancer stage has an important role in choosing the treatment, and a patient’s response to treatment depends on her age, so we divided the dataset according to the age of patient at diagnosis of disease (age < 50, age ≥ 50) and the stage of cancer (I, II, III, IV). This produced eight combinations of age and stage with different conditions of PH and AFT assumptions, and we could compare the treatment effect on patient survival with parametric and semi-parametric models.

The presence of hormone receptors has been proven to have an effect on survival time of patient, and was included in all models.

For each combination of age and stage, we only included variables for which more than 10 patients received and did not receive the treatment.

**RESULTS**

Frequency of patients receiving treatment by stages at diagnosis is shown in Table 1. Results of comparison between parametric regression models have been summarized in Table 2. Also, the results of fitting parametric model and Cox model are shown in Tables 3 and 4 respectively.

**Situations when both PH and AFT assumptions are satisfied**

For patients under age 50 years in disease stages I and IV, the PH and AFT assumptions were satisfied.

In patients under age 50 years with stage I cancer, the best-fitting parametric model used a conventional GG distribution in which chemotherapy \((P < 0.001)\) and erposneg \((P = 0.002)\) were significant. For the Cox model, radiotherapy \((P = 0.001)\), chemotherapy \((P < 0.001)\), and hormone therapy \((P = 0.01)\) were significant. In both models, no interactions were significant.

In patients under age 50 years with stage IV cancer, a conventional lognormal model was the best candidate in GG family. The AIC and BIC criteria were the same for conventional lognormal and conventional log-logistic models, but the Cox–Snell residual plot indicated better fit for the lognormal. Radiotherapy \((P = 0.01)\), hormone therapy \((P = 0.01)\), and erposneg\(\) were meaningful in lognormal model. For the Cox model, radiotherapy \((P = 0.02)\) and erposneg \((P = 0.001)\) were significant and no interactions
Table 1: Number of patients receiving treatment by stage at diagnosis

| Stage | Treatment      | Age <50 | No | Age ≥50 | No |
|-------|----------------|---------|----|---------|----|
| I     | Surgery        | 1217    | 2  | 4343    | 12 |
|       | Radiotherapy   | 955     | 264| 2976    | 1379 |
|       | Chemotherapy   | 443     | 776| 229     | 4126 |
|       | Hormone therapy| 233     | 986| 1546    | 2809 |
| II    | Surgery        | 1815    | 6  | 3748    | 12 |
|       | Radiotherapy   | 1486    | 335| 2547    | 1213 |
|       | Chemotherapy   | 1676    | 145| 1169    | 2591 |
|       | Hormone therapy| 726     | 1095| 2955   | 805 |
| III   | Surgery        | 288     | 17 | 538     | 58 |
|       | Radiotherapy   | 293     | 12 | 509     | 87 |
|       | Chemotherapy   | 298     | 7  | 326     | 270 |
|       | Hormone therapy| 165     | 140| 464     | 132 |
| IV    | Surgery        | 73      | 33 | 228     | 141 |
|       | Radiotherapy   | 87      | 19 | 266     | 103 |
|       | Chemotherapy   | 86      | 20 | 128     | 241 |
|       | Hormone therapy| 65      | 41 | 294     | 75 |

Table 2: Comparison between parametric regression models in each category of age and stage

| Stage | Distribution* | Age < 50 |          |          | Age ≥50 |          |          |
|-------|---------------|----------|----------|----------|---------|----------|----------|
|       |               | XL       | BIC      | AIC      | XL       | BIC      | AIC      |
| I     | GGS           | -------- | -------- | -------- | -------- | -------- | -------- |
|       | GGe           | Ref      | -561.24  | -550.37  | Ref      | -1587.83 | -1572.5  |
|       | Weibull       | 26.82    | -571.1   | -561.78  | 36.4294  | -1601.86 | -1588.72 |
|       | Log-normal    | 11.5     | -563.44  | -554.12  | 6.9262   | -1587.11 | -1573.97 |
|       | Log-logistic  | -569.297 | -559.98  | -1598.5  | -1585.38 |
| II    | GGS           | Ref      | -1548.12 | -1527.1  | Ref      | -2998.45 | -2979.41 |
|       | GGe           | 131.96   | -1591.58 | -1581.06 | 91.6     | -3027.8  | -3017.21 |
|       | Weibull       | 200.57   | -1633.39 | -1619.36 | 257.56   | -3114.89 | -3102.19 |
|       | Log-normal    | 116.31   | -1591.26 | -1577.23 | 123.56   | -3047.89 | -3035.19 |
|       | Log-logistic  | -1616.24 | -1602.21 | -3087    | -3074.3  |
| III   | GGS           | -------- | -------- | -------- | Ref      | -757.44  | -735.93  |
|       | GGe           | Ref      | -376.13  | -371.83  | 29.03    | -739.82  | -730.44  |
|       | Weibull       | -91.2    | -418.96  | -415.42  | 107.89   | -792.21  | -777.869 |
|       | Log-normal    | 52.6     | -399.56  | -396.12  | 48.93    | -762.74  | -748.4   |
|       | Log-logistic  | -406.88  | -403.44  | -771.4   | -757.06  |
| IV    | GGS           | -------- | -------- | -------- | Ref      | -570.95  | -556.62  |
|       | GGe           | Ref      | -155.34  | -152.69  | 23.32    | -558.97  | -552.28  |
|       | Weibull       | 26.36    | -166.18  | -163.86  | 59.2     | -585.77  | -576.22  |
|       | Log-normal    | 4.1      | -155.05  | -152.73  | 10.06    | -561.2   | -541.65  |
|       | Log-logistic  | -153.76  | -151.44  | -560.485 | -550.93  |

X_L: Likelihood ratio statistic relative to generalized gamma , Ref: Reference distribution, ----: not meaningful, s: saturated (all parameters depend on covariates and model is not AFT), c: conventional ( only location parameter depends on covariates and model is AFT), *: in each category of age and stage, all distributions’ conditions are like the reference distribution of that category whether it is saturated or conventional.

were significant in any of the models. Cox–Snell residual plots for the lognormal, log-logistic, and Cox PH model are given in Figures 1–3.

Situations when only AFT assumptions hold

For patients aged 50 years or more with stage I cancer, the PH assumption was not satisfied, but the
Table 3: Relative time estimated in parametric models

| Stage | Variable   | Age <50          | Age ≥50          |
|-------|------------|------------------|------------------|
|       |            | RT   | 95%CI      | RT   | 95%CI      |
| I     | Radiotherapy | ----- | -----       | Radiotherapy | 0.772 | (0.63,0.95) |
|       | Chemotherapy | 0.547 | (0.39,0.76) | Chemotherapy | 0.405 | (0.28,0.58) |
|       | Hormone therapy | ----- | -----       | Hormone therapy | 0.553 | (0.46,0.67) |
|       | Erposneg    | 1.71  | (1.22,2.41) | Erposneg    | 1.621 | (1.27,2.05) |
| II    | Radiotherapy | NC    | NC          | Radiotherapy | NC    | NC          |
|       | Hormone therapy | NC    | NC          | Erposneg    | NC    | NC          |
|       | Erposneg    | NC    | NC          | Erposneg    | NC    | NC          |
| III   | Hormone therapy | 1.356 | (1.05,1.75) | Surgery     | NC    | NC          |
|       | Erposneg    | 1.531 | (1.165,2.01) | Radiotherapy | NC    | NC          |
|       | Chemotherapy | 0.547 | (0.39,0.76) | Chemotherapy | NC    | NC          |
|       | Hormone therapy | 1.77  | (1.13,2.8)  | Hormone therapy | NC    | NC          |
|       | Erposneg    | 1.82  | (1.14,2.9)  | Erposneg    | NC    | NC          |

Table 4: Relative hazard estimated in Cox model

| Stage | Variable   | Age <50          | Age ≥50          |
|-------|------------|------------------|------------------|
|       |            | HR  | 95%CI      | HR  | 95%CI      |
| I     | Radiotherapy | 2.53  | (1.45,4.39) | Radiotherapy | 1.42  | (1.13,1.79) |
|       | Chemotherapy | 1.9   | (1.42,2.8)  | Chemotherapy | 2.39  | (1.77,3.22) |
|       | Hormone therapy | 1.66  | (1.12,2.46) | Hormone therapy | 1.91  | (1.56,2.34) |
|       | Erposneg    | ----- | -----       | Erposneg    | -PH,s | -----       |
| II    | Radiotherapy | -PH   | -----       | Radiotherapy | 2.21  | (1.88,2.59) |
|       | Chemotherapy | ----- | -----       | Chemotherapy | ----- | -----       |
|       | Hormone therapy | -PH   | -----       | Hormone therapy | ----- | -----       |
|       | Erposneg    | -PH   | -----       | Erposneg    | -PH,s | -----       |
| III   | Hormone therapy | ----- | -----       | Surgery     | 0.49  | (0.35,0.68) |
|       | Erposneg    | ----- | -----       | Radiotherapy | ----- | -----       |
|       | Chemotherapy | ----- | -----       | Chemotherapy | ----- | -----       |
|       | Hormone therapy | ----- | -----       | Hormone therapy | ----- | -----       |
|       | Erposneg    | -PH   | -----       | Erposneg    | -PH,s | -----       |
| IV    | Surgery     | ----- | -----       | Surgery     | 0.66  | (0.53,0.84) |
|       | Radiotherapy | 0.50  | (0.3,0.9)   | Radiotherapy | -PH   | -----       |
|       | Chemotherapy | ----- | -----       | Chemotherapy | ----- | -----       |
|       | Hormone therapy | ----- | -----       | Hormone therapy | -PH   | -----       |
|       | Erposneg    | 0.42  | (0.25,0.7)  | Erposneg    | -PH,s | -----       |

RT: Relative time       -----: Not significant       NC: Not constant

HR: Relative hazard       -----: Not significant       -PH: Not PH       S: Stratified by

AFT assumptions held. In parametric models, the conventional GG had the best fit and radiotherapy ($P = 0.014$), chemotherapy ($P < 0.001$), hormone therapy ($P < 0.001$), and erposneg ($P < 0.001$) were significant. In semiparametric model, the covariates radiotherapy ($P = 0.003$), chemotherapy ($P < 0.001$), hormone therapy ($P < 0.001$), and erposneg ($P < 0.001$) were meaningful; and since the variable erposneg did not hold the PH assumption, it was lost due to stratification and
the effects of other parameters were estimated by stratifying the model by this variable. No interactions were significant in each model.

**Situations where none of the PH and AFT assumptions are satisfied**

For patients aged 50 years or more with stage IV cancer, none of the PH and AFT assumptions held. In parametric model, the GG family fitted better than the log-logistic according to AIC and BIC criteria. In models based on the GG family, the saturated lognormal distribution was the best model. In this model, surgery \((P = 0.001)\), radiotherapy \((P = 0.004)\), hormone therapy \((P < 0.001)\), and erposneg \((P < 0.001)\) were significant and the interaction of hormone therapy and erposneg \((P = 0.021)\) was meaningful. Since the saturated lognormal is not an AFT model, the relative survival time was calculated for the 25\(^{th}\), 50\(^{th}\), and 75\(^{th}\) quantiles. Results of fitting have been shown in Table 5. In the Cox model, surgery \((P = 0.001)\), radiotherapy \((P = 0.03)\), hormone therapy \((P < 0.001)\), and erposneg \((P = 0.001)\) were significant. Since the PH assumption only held for surgery, the effect of this variable was estimated by stratifying on erposneg. No interaction was significant in each model.

For all patients with stage II cancer, and patients aged 50 and older with stage III cancer, neither the PH nor the AFT assumptions were satisfied. In semiparametric models, most variables did not meet the PH assumption and their effect could not be estimated; in parametric models, the saturated GG fitted better than other distributions.

**DISCUSSION**

When PH and AFT assumptions were satisfied, both the parametric and semiparametric models were appropriate. The models indicate different significant variables, but parametric models have some advantages over semiparametric models in general. With small sample sizes, relative efficiencies may further change in favor of parametric models.\(^{[4]}\) When the PH assumption is satisfied, some studies indicate that parametric PH models are a better approach than the Cox model.\(^{[6,7]}\) Further, some studies have shown the robustness of parametric AFT models to misspecification because of their log-linear form.\(^{[8]}\) Finally, one advantage of a parametric model compared to a Cox model is that the parametric likelihood accommodates right-,
Table 5: Relative times for stage IV–Age ≥ 50

| Variable                        | RT(0.25) | RT(0.5) | RT(0.75) |
|---------------------------------|----------|---------|----------|
| Surgery                         | 1.3      | 1.55    | 1.86     |
| Radiotherapy                    | 2.09     | 1.68    | 1.35     |
| Erospneg                        | 2.84     | 3       | 3.17     |
| Hormone therapy if erosneg=1    | 2.4      | 2.53    | 2.67     |
| Hormone therapy if erosneg=0    | 1.14     | 1.19    | 1.24     |

left-, or interval-censored data. The Cox likelihood, by contrast, handles right-censored data but does not directly accommodate left- or interval-censored data.\cite{2}

When the PH assumption is not satisfied but AFT assumptions hold, the parametric model can be used as a substitute for the Cox model. Other studies have suggested the same thing.\cite{9, 10}

When neither the PH nor the AFT assumptions hold, a member of the GG distribution, the saturated lognormal, can be used to calculate relative survival times in different quantiles. The lognormal distribution has a long history in cancer survival analysis.\cite{10, 11}. In many settings, including breast cancer analysis, where the proportionality assumption does not hold, the lognormal model has been shown to be appropriate.\cite{12-14}. A meta-analysis has shown that saturated parametric models provide better results than conventional models for comparing treatments.\cite{15}

When the PH and AFT assumptions do not hold and the saturated GG distribution fits better than other distributions, further analyses may need to be considered. Through its three parameters, the GG family contains many different distributions such as the inverse Weibull and inverse lognormal.\cite{3} Accordingly, the best fit can be found by trying different parameters. Our analysis applied the most commonly used parametric distributions in survival analysis, but we could not determine the best fit.

For one category of age and stage in our study, according to AIC and BIC criteria, the log-logistic distribution gave the same fit as the lognormal model within the GG family. The log-logistic distribution belongs to the generalized F distribution, which includes the GG distribution. When a member of the GG family does not fit satisfactorily, the best distribution can be found through the generalized F distribution family.\cite{16, 17}

If the scales of the parameters in Cox’s model and the parametric models differ, neither parameter estimates nor their estimated variances are suitable for comparisons. If the Cox model is compared to parametric PH models, the efficiency of parameter estimates can be compared by Wald-type statistics.\cite{4}

For patients with stage IV cancer, semiparametric and parametric models showed lower hazards and longer survival times for patients receiving treatments than those not receiving treatments. For patients with stage I cancer in both the under 50-year and over 50-year age groups, semiparametric and parametric models showed higher hazards and smaller survival times for patients receiving treatments, which might reflect other covariates that caused the patients to receive a certain treatment. For example, studies have shown the ethnicity of patients affect their use of treatments that are more common in lower stages of cancer.\cite{18}

In both parametric and semiparametric models applied to different combinations of age and cancer stage, the expression of hormone receptors was associated with a longer survival time and lower hazard, which has been confirmed in many other studies.\cite{19}

A major strength of this study is that we fitted models and performed comparisons using a large set of real-life data from a population-based cancer registry. The major limitation is that our findings describe associations with survival, and not causes of survival. In particular, breast cancer patients often receive treatments because of disease characteristics. A patient’s survival does not necessarily depend on the treatment they receive; rather, the treatment that a patient receives often depends on disease characteristics that predict survival.

Subsequent research should examine whether our findings hold for other diseases and other populations. In particular, it would be of interest to determine whether the findings are sustained in more-recent patient cohorts – where newer treatments have been used. However, there are
some recent studies that describe the factors associated with survival time of patients in developing countries.[20, 21]

**CONCLUSION**

When PH and AFT assumptions were satisfied, semiparametric and parametric models provided two different valid approaches for exploring breast cancer patients’ survival, and the models can be seen as complementary. When PH assumptions were not satisfied but AFT conditions held, the parametric model should be used instead of the Cox model. When neither the PH nor the AFT assumptions were met, the log normal distribution, a member of the GG family, provided an alternative approach to semiparametric model. More generally, when PH assumptions are not satisfied, parametric models should be considered, whether or not AFT assumptions are met.

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