Role of time scale in the spreading of asymmetrically interacting diseases

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Diseases and other contagion phenomena in nature and society can interact asymmetrically, such that one can benefit from the other, which in turn impairs the first, in analogy with predator-prey systems. Here, we consider two models for interacting disease-like dynamics with asymmetric interactions and different associated time scales. Using rate equations for homogeneously mixed populations, we show that the stationary prevalences and phase diagrams of each model behave differently with respect to variations of the relative time scales. We also characterize in detail the regime where transient oscillations are observed, a pattern that is inherent to asymmetrical interactions but often ignored in the literature. Our results contribute to a better understanding of disease dynamics in particular, and interacting processes in general, and could provide interesting insights for real-world applications.

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I. INTRODUCTION

Spreading processes are ubiquitous in nature and society. A primary example of these phenomena is the propagation of diseases in human and animal populations [1,2]. Despite many recent advances in the theoretical, computational, and data-driven modeling of diseases, there are still many scientific problems that remain open. Two such problems are currently of high relevance. On the one hand, we have many limitations when it comes to understanding the effects of mobility restrictions and human behavioral changes on the evolution of a disease [3]. On the other hand, there are many diseases that rarely evolve in isolation; on the contrary, diverse viral strains or different pathogens can either compete for the susceptible population or establish cooperative interactions, both at a population level and inside a host organism [4]. This paper focuses on studying spreading processes with the aim of shedding some light on the second kind of challenge.

Early works on interacting diseases [5–9] have modeled the dynamics of pathogens or strains that interact competitively, showing that their coexistence is not always possible. In the last years, these models have been adapted to networks [10–16], including single-layer and multiplex networks [2,17], as well as metapopulations [14]. These works have shown that the network organization plays a fundamental role in the evolution of the dynamical processes, affecting the epidemic threshold and prevalence (e.g., Ref. [18]). Recently, the focus has also been placed on collaborative contagion, in which there is a positive feedback between diseases [19–26]. In these models, discontinuous phase transitions, in which the prevalence goes from zero to a finite value abruptly, have been observed, as well as simultaneous stability of two epidemic states. It is worth remarking that although the network topology can generate important phenomena, some essential properties can still be observed in homogeneously mixed populations, as demonstrated by Chen et al. [25].

General models for interacting diseases have also been developed recently. This is the case in Ref. [27], where the authors introduced a generic model for two interacting diseases in multiplex networks, comprising the cases of competitive, collaborative, and asymmetrical interactions, for both susceptible-infected-susceptible (SIS) and susceptible-infected-removed (SIR) compartmental models. They calculated the respective epidemic thresholds and studied the prevalence for both the competitive and the collaborative scenarios. More recently, this model inspired a Markov chain approach [28]. In all works, however, the asymmetrical case has remained unexplored.

Asymmetrical interactions occur in many complex systems. For example, there are reported cases of HIV viral load suppression by the presence of some other pathogens [29–31], which suggests an asymmetrical interaction between HIV and many diseases. Asymmetrical interactions have been studied mostly in the case of the interplay between epidemic and awareness [18,32–37], malicious computer worms and spreading countermeasures [38–41], and antigens and immune system agents [42–44], among other examples. In the case of disease and awareness, the epidemic stimulates
information awareness, which in turn tends to reduce the exposure to disease and therefore also disease prevalence, configuring an asymmetrical interaction between the two processes. It has been shown that, in general, the awareness can effectively reduce the disease prevalence and increase its threshold [32,33]. The epidemic, in turn, can sustain an information outbreak even when the latter is below its “independent” threshold [18,34]. Correlations and structural properties of the underlying multiplex network can also influence this propagation process [18,32,34].

As mentioned above, some previous works have examined asymmetrically interacting spreading phenomena. However, this has been done for specific cases, and a general description of these processes is, to the best of our knowledge, still missing. Of special interest is the fact that the interacting dynamics can do so at different time scales, whose effect has not been deeply investigated yet. Changes in the relative time scale (i.e., the possibility that one process has a different intrinsic clock concerning the other) typically do not change the qualitative behavior of interacting systems but often lead to relevant quantitative effects. For example, Karrer and Newman [11] showed that for competing SIR diseases, the slower pathogen may take advantage and even win the competition due to its inherent time scale. Oliveira and Dickman [45] also showed that the slower species can win a competition in a contact process (CP), including scenarios with periodic and stochastic environmental variations. In the context of asymmetrical interactions, Wu et al. [44] and Poletto et al. [46] demonstrated that the epidemic threshold is affected by changes in the recovery rate when the reproductive number (the ratio between the spreading and the recovery rates) is kept constant, which is equivalent to a variation of the time scale. Equivalently, it has also been shown that a slower information awareness may have more impact on the epidemic prevalence than a faster one [36,37]. Therefore the development of a more general understanding of the influence of the relative time scales of asymmetrically interacting disease-like models is of great interest.

In this paper, we study two SIS-like models for interacting diseases when they are in an asymmetrically interacting regime over homogeneously mixed populations and continuous-time evolution. We rely on the most simple setup that does not involve structured populations to extract the intrinsic properties of the models. In each model, we assign a parameter \( \pi \) that allows us to control the relative time scale between the two dynamics while keeping constant the epidemic forces of each disease. We adopt a descriptive analysis of each model, determining whether different aspects are or are not influenced by the time scale, including the respective phase diagrams, coexistence, and oscillatory behavior. We also compare both models and discuss our results in the light of some previous works.

II. EPIDEMIC MODELS WITH ASYMMETRICAL INTERACTION

In what follows, regardless of the model considered, we assume that there are two diseases that interact. Moreover, we consider that disease I is the “prey” (it is impaired by the other disease) and disease II is the “predator” (it is benefited from the presence of the first disease), in an analogy with the asymmetrical interaction of predator-prey systems. A parameter \( \pi \) controls the relative time scale between the two diseases. It is implemented as follows: The rates of all processes promoted by disease I are multiplied by \((1 - \pi)\), whereas the rates of disease II processes are multiplied by \(\pi\). Making \( \pi \) range between 0 and 1, we sweep through scenarios in which disease I is faster \((\pi < 0.5)\) or slower \((\pi > 0.5)\) than disease II, as well as the balanced case \((\pi = 0.5)\). Although \( \pi \) does not add a new degree of freedom to the model, the advantages of using this approach are as follows: (i) It allows us to control the relative time scale with a single parameter; (ii) the “overall rate” of the system, which can be regarded as \(\sim([1 - \pi] + \pi)\), is kept constant when varying \( \pi \); and (iii) plotting variables as a function of \( \pi \) is simple because it is limited between 0 and 1. Let us now describe each of the models scrutinized in the rest of the paper.

A. Model A: Interacting diseases through susceptibility change

In this variant, we consider that the presence of one disease impairs the spreading of the second one by changing the individual’s susceptibility to catch the other disease. It has been used to describe the dynamics of competing pathogens with partial cross immunity [9] and for collaborative contagion [25,27], but its asymmetrical version is still largely unexplored. The latter regime is a prototypical model to describe processes in which an epidemic coexists with an information dynamics in which awareness regarding the disease plays a role in the spreading of it. Indeed, our model A is mathematically similar to models used for that purpose [37].

In this model, each individual can either be susceptible to both diseases \((S_1S_2)\), infected by one disease and susceptible to the second one \((I_1S_2)\) or \((S_1I_2)\), or infected by both diseases \((I_1I_2)\). Note that we denote as \(I_1\) and \(I_2\) individuals that are infected by disease I and II, respectively, regardless of their state with respect to the other disease. For completely susceptible individuals \(S_1S_2\), the baseline contagion rate of disease I (II) when in contact with an individual infected by disease I (II) is \(\beta_1 (\beta_2)\). For individuals already infected by disease I but susceptible to disease II \((I_1S_2)\), the contagion rate for disease II is \(\Gamma_2 \cdot \beta_2\); that is, it is multiplied by a factor \(\Gamma_2\). The same holds for \(S_1I_2\) individuals, for which the contagion rate by disease I is changed to \(\Gamma_1 \cdot \beta_1\). The healing rates from disease I and II are \(\mu_1\) and \(\mu_2\), respectively, and are not affected by the other disease. We also define \(\lambda_1 = \beta_1/\mu_1\) and \(\lambda_2 = \beta_2/\mu_2\), which are the basic reproduction numbers of each disease as if they were independent. Figure 1 represents all the possible transitions for this model, with their respective time scale factors as explained before.

The asymmetrical interaction between the two diseases can be obtained by setting \(0 \leq \Gamma_1 < 1\) and \(\Gamma_2 > 1\). This means that individuals that hold disease II are less susceptible to catch disease I in comparison with fully susceptible individuals. On the other hand, individuals infected by disease I are more likely to catch disease II. Therefore disease I enhances the propagation of disease II, whereas disease II impairs the propagation of disease I. We represent the density of individuals in a given state \(X\) by \(\rho_X\), with \(X\) being either a composite state (such as \(I_1S_2\)) or a simple state (such as \(I_1\)). Based on the diagram from Fig. 1, the time evolution of the composite state
FIG. 1. State transitions allowed in model A. The baseline infection and healing rates of disease I (II) are $\beta_1 (\beta_2)$ and $\mu_1 (\mu_2)$, respectively, $\Gamma_1 (\Gamma_2)$ represents the modification to the baseline transmission rate of disease I (II) due to the presence of the other disease in the host. In addition, each rate is multiplied by its corresponding time scale factor: $(1 - \pi)$ for processes of disease I and $\pi$ for disease II.

The densities are yet subject to the normalization constraint $\rho_{i12} + \rho_{i1i} + \rho_{i2i} + \rho_{i2i} = 1$. This constraint makes the system effectively three dimensional. One can reduce the number of equations and simplify the notation by using the following variable change (as done in Ref. [25]):

$$u = \rho_{i1} = \rho_{i12} + \rho_{i1i},$$
$$v = \rho_{i2} = \rho_{i2i} + \rho_{i2i},$$
$$w = \rho_{i2i},$$

for which the dynamical equations are

$$\dot{u} = (1 - \pi)(\beta_1(1 - u) + (\Gamma_1 - 1)\beta_1(v - w) - \mu_1)u,$$  \hspace{0.5cm}(1)

$$\dot{v} = \pi[\beta_2(1 - v) + (\Gamma_2 - 1)\beta_2(u - w) - \mu_2]v,$$  \hspace{0.5cm}(2)

$$\dot{w} = (1 - \pi)(\Gamma_1\beta_1(v - w)u - \mu_1)w + \pi[\Gamma_2\beta_2(u - w)v - \mu_2]w.$$  \hspace{0.5cm}(3)

The densities is given by the following equations:

$$\frac{d\rho_{i12}}{dt} = -(1 - \pi)\beta_1\rho_{i12}\rho_{i1} - \pi\beta_2\rho_{i12}\rho_{i2} + (1 - \pi)\mu_1\rho_{i12} + \pi\mu_2\rho_{i12},$$

$$\frac{d\rho_{i1i}}{dt} = -(1 - \pi)\Gamma_1\beta_1\rho_{i1i}\rho_{i1} + \pi\beta_2\rho_{i1i}\rho_{i2} + (1 - \pi)\mu_1\rho_{i1i} - \pi\mu_2\rho_{i1i},$$

$$\frac{d\rho_{i2i}}{dt} = -(1 - \pi)\Gamma_2\beta_2\rho_{i2i}\rho_{i2} - \pi\Gamma_2\beta_2\rho_{i2i}\rho_{i2} - (1 - \pi)\mu_1\rho_{i2i} + \pi\mu_2\rho_{i2i},$$

$$\frac{d\rho_{i2i}}{dt} = -(1 - \pi)\Gamma_1\beta_1\rho_{i2i}\rho_{i1} + \pi\Gamma_2\beta_2\rho_{i2i}\rho_{i2} - (1 - \pi)\mu_1\rho_{i2i} - \pi\mu_2\rho_{i2i}.$$

The model is schematically represented in Fig. 1. Despite the competitive interaction between computer viruses and the spreading of antimalware, or the dynamics of fake news and fact-checking messages. In such scenarios, either the antimalware or the fact-checked message replaces the previous "infectious agent." We also note that our model is a particular case of the model discussed in Ref. [44] and can also be interpreted as a generalized predator-prey model [55].

Specifically, in model B, we represent susceptible individuals simply by $S$ and infected individuals of diseases I and II by $I_1$ and $I_2$, respectively. The transmission rates are $\beta_1$ and $\beta_2$, the healing rates are $\mu_1$ and $\mu_2$, and the rate at which $I_1$ individuals are "superinfected" by disease II when exposed to $I_2$ individuals is given by a modified term $\alpha \cdot \beta_2$. As in model A, each term is also multiplied by the corresponding time scale factor $(1 - \pi)$ for disease I and $\pi$ for disease II, and also $\lambda_1 = \beta_1 / \mu_1$ and $\lambda_2 = \beta_2 / \mu_2$ are the basic reproduction numbers of the independent diseases. The transitions are schematically represented in Fig. 2.
aspect of the model, one can generate an asymmetrical interaction by setting $\alpha > 1$. This is because $I_1$ individuals, in this case, are more easily infected by disease II than susceptible ones, meaning that the spreading of disease II is enhanced by the presence of disease I, which in turn is at a disadvantage due to the competition with disease II.

Using the same notation as before, $\rho_i$, for the density of individuals in state $X$, we write the dynamical equations

$$\frac{d\rho_i}{dt} = -(1 - \pi)\beta_1 \rho_i \rho_i - \pi \beta_2 \rho_i \rho_i + (1 - \pi)\mu_1 \rho_i + \pi \mu_2 \rho_i,$$

$$\frac{d\rho_i}{dt} = (1 - \pi)\beta_1 \rho_i \rho_i - \pi \alpha \beta_2 \rho_i \rho_i - (1 - \pi)\mu_1 \rho_i,$$

$$\frac{d\rho_i}{dt} = \pi \beta_2 \rho_i \rho_i - \pi \alpha \beta_2 \rho_i \rho_i - \pi \mu_2 \rho_i,$$

where the normalization constraint is $\rho_x + \rho_1 + \rho_2 = 1$. As in model A, we make the change of variables $u = \rho_{1,2}$, $v = \rho_i$, to obtain the reduced set of dynamical equations

$$\dot{u} = (1 - \pi)\beta_1 (1 - u - v) - \mu_1 u - \pi \alpha \beta_2 uv,$$

$$\dot{v} = \pi \beta_2 (1 - u - v) - \mu_2 v + \pi \alpha \beta_2 uv.$$

### III. RESULTS

#### A. Phase diagrams

In the asymmetrically interacting regime, models A and B share a common feature: For any given set of parameters, there is exactly one stable fixed point within the region of the phase portrait that corresponds to physically possible solutions. Therefore, unlike what has been reported for mutually competitive or cooperative scenarios [25,51], our models do not present bistability [56]. This is intrinsic to the positive-negative feedback of the asymmetric interaction between the diseases. Both models A and B present four phases, separated by transcritical bifurcations: P1, no disease; P2, disease I only; P3, disease II only; and P4, coexistence of both diseases. Here, we investigate the $\lambda_1 \times \lambda_2$ phase diagrams, setting the other parameters to fixed values. Figure 3 shows the phase diagrams for both models.

Through a stability analysis, we can obtain the phase transition curves of both models. The boundary between phases P2 and P4 in model A is expressed as

$$\frac{1}{\lambda_2} = 1 + (\Gamma_2 - 1) \left(1 - \frac{1}{\lambda_1}\right), \quad \lambda_1 > 1,$$

whereas, by the symmetry of the model, the boundary between phases P3 and P4 is given by

$$\frac{1}{\lambda_1} = 1 - (1 - \Gamma_1) \left(1 - \frac{1}{\lambda_2}\right), \quad \lambda_2 > 1.$$

From Eq. (7), we can check that when $\Gamma_1 = 1$ (i.e., no interaction), the boundary between phases P3 and P4 is a vertical line. On the other limit, for $\Gamma_1 = 0$, the boundary becomes the identity line $\lambda_1 = \lambda_2$.

The other two bifurcations are trivial and are given by $\lambda_2 = 1$ for $0 \leq \lambda_1 \leq 1$ (boundary between P1 and P3) and $\lambda_1 = 1$ for $0 \leq \lambda_2 \leq 1$ (boundary between P1 and P2). It is worth noticing that for model A, none of the phase transition curves depends on the time scale parameter $\pi$.

The phase transition curves have similar shapes to those reported in other works using single-layer and multiplex [27] networks, which shows that such feature is essential to the model itself. It can also be shown that the transition curves between P1 and P2 and between P3 and P4 are equivalent to the curve between healthy and endemic phases in a model for epidemics with awareness [32].

For model B, the boundary between P2 and P4 is given by

$$\frac{1}{\lambda_2} = 1 + (\alpha - 1) \left(1 - \frac{1}{\lambda_1}\right), \quad \lambda_1 > 1,$$
which is similar to that in Eq. (6), only replacing $\Gamma_2$ by $\alpha$. The boundary between P3 and P4 is expressed as

$$\lambda_1 = [\alpha \chi (\lambda_2 - 1) + 1] \lambda_2, \quad \lambda_2 > 1, \quad (9)$$

where we define $\chi$ (interpreted as the time scale ratio between diseases II and I) as

$$\chi = \frac{\pi \mu_2}{(1 - \pi) \mu_1}.$$ 

The trivial boundaries of region P1 with regions P2 and P3 are the same as in model A. Notice, however, that the parameter $\chi$ depends increasingly on the time scale parameter $\pi$, and so does the critical $\lambda_1$ value from Eq. (9). This means that as $\pi$ increases (i.e., when disease II propagates on shorter time scales), the phase transition between P3 and P4 moves to lower values of $\lambda_2$, contracting the region of coexistence and approaching the horizontal line $\lambda_2 = 1$ as $\pi \to 1$ (notice that $\chi$ diverges). Therefore, in this model, a faster clock for disease II makes it more effective to suppress disease I. This is an interesting result that might be used to control the prevalence of disease (or any other “infectious agent”) in the host population. In the limit of $\pi \to 0$, the transition between P3 and P4 becomes the identity line $\lambda_1 = \lambda_2$. The $\pi$ dependence of the phase diagrams of model B constitutes an important feature that depends on the specific shape of the considered model.

The reason for such difference in model B is the existence of superinfection, which is a process of disease II that modifies the state of an individual for both diseases. In contrast, for model A, the processes of a given disease can only change its own states, making the relative time scale irrelevant near the phase transitions.

### B. Behavior of the stationary prevalence

In the phase of coexistence (region P4 of the phase diagrams), both models present a single stable fixed point, for which $u, v \neq 0$. One can use the dynamical equations [Eqs. (1)–(3) for model A and Eqs. (4) and (5) for model B] to derive analytical expressions for the prevalences at the coexistence fixed point. The derivation and the final expressions are shown in Appendix A. As expected, the stationary prevalence of each disease is a nondecreasing function of its reproduction ratio $\lambda = \beta/\mu$, when considering other parameters as fixed. However, the dependence of the prevalences with the relative time scale parameter $\pi$ is not trivial and is different in each model.

Figure 4 shows the basic behavior of the fixed point prevalences with $\pi$ for models A and B. While the prevalence $v$ of disease II decreases with $\pi$ for both models, the prevalence $u$ of disease I has opposite behaviors in each of them. In model A, the prevalence $u$ of coinfection increases with $\pi$. This variant shows the same behavior (if equating disease II with the information) as recently reported in some works on epidemics with awareness [36,37] in complex networks; namely, a faster relative clock of the information induces an increase in the disease prevalence. For model B, however, the behavior is the opposite: The prevalence of disease I decreases with $\pi$. Considering also how the time scale parameter distorts the phase diagram of model B [see Fig. 3, where the $\lambda_1, \lambda_2$ values used in Fig. 4 are indicated by dotted lines], we see that a faster clock of the “predator” process (disease II) effectively decreases the spreading of the “prey” process, leading to its extinction if $\lambda_2 > 1$ and for sufficiently large $\pi$. Therefore the relationship between the prevalences and the relative time scale in asymmetrically interacting spreading phenomena is a feature that depends on the specific shape of the considered model.

In Figs. 5 and 6, we further analyze the behavior of the prevalences with $\pi$ and the parameters that control the interactions: $\Gamma_1$, $\Gamma_2$ for model A and $\alpha$ for model B. We first notice that the increasing or decreasing trends with respect
to \( \pi \), as observed in Fig. 4, are not changed for different values of the interaction parameters: Disease I increases while disease II decreases with \( \pi \) for model A (Fig. 5), whereas both prevalences decrease with \( \pi \) for model B (Fig. 6). To rule out the possibility that there is a region of the parameter space in which the reported behaviors with respect to \( \pi \) might be different, we show in Appendix B that this is not possible, i.e., that the behavior with \( \pi \) is always the same for each model in the coexistence region.

We can also analyze how the prevalences vary with changes in the interaction parameters. For model A (Fig. 5), we see that an increase in \( \Gamma_1 \) increases both the prevalences of disease I [Fig. 5(a)] and II [Fig. 5(b)]. This is expected, as a greater \( \Gamma_1 \) value means a weaker impairing of the propagation of disease I, which is beneficial for both diseases (as disease II benefits from disease I). On the other hand, an increase in \( \Gamma_2 \) causes a decrease in disease I [Fig. 5(c)] and an increase in disease II [Fig. 5(d)]. This is also expected, as a larger value of \( \Gamma_2 \) means a greater benefit to disease II, which in turn is detrimental to disease I. Therefore, for model A, the effect of the two interaction parameters in each disease’s prevalence is intuitive and predictable. The effects are also numerically influenced by the time scale \( \pi \), yet not qualitatively changed.

However, for model B, which has a single interaction parameter \( \alpha \), the behavior of the prevalence is not trivial. From Fig. 6, we see that while disease I prevalence [Fig. 6(a)] is always reduced with an increase in \( \alpha \), the behavior of the prevalence of disease II [Fig. 6(b)] with \( \alpha \) is not uniform and may have an optimal value that depends on \( \pi \). This happens because the superinfection transition, controlled by \( \alpha \), is simultaneously beneficial to disease II and detrimental to disease I. Thus, as seen from model A, an increase in \( \alpha \) certainly reduces the prevalence of disease I but has a “conflicting” effect on the prevalence of disease II. While a value of \( \alpha \) close to 1 means almost no benefit to disease II from disease I, a large value \( \alpha \gg 1 \) means an excessive “predation” from disease II; therefore there exists an optimal intensity of the interaction \( \alpha \). This is further illustrated in Fig. 7, where we show the prevalence of disease II in model B as a function of \( \alpha \), for different values of \( \pi \). As can be seen, there is an optimal value of \( \alpha \) that maximizes the prevalence, for fixed values of the other parameters.

FIG. 6. Stationary prevalences of (a) disease I and (b) disease II for model B, plotted as functions of \( \pi \) and \( \alpha \). The reproduction ratios are set to \( \lambda_1 = 1.7 \) and \( \lambda_2 = 1.2 \).

FIG. 7. Stationary prevalences of disease II for model B, plotted as a function of \( \alpha \) and for different values of \( \pi \). The dashed line represents the optimal value of \( \alpha \) and its corresponding \( \pi \) as \( \pi \) is continuously changed. The other parameters are set to \( \beta_1 = 1.7 \), \( \beta_2 = 1.2 \) and \( \mu_1 = \mu_2 = 1 \).

C. Damped oscillations: Node versus spiral point

Some predator-prey systems, which naturally have an asymmetrical relationship between two processes, are known to present stable closed orbits (i.e., sustained oscillations) [55]. For both epidemic models addressed in our work, there are no closed orbits in the physical region of the phase portrait [57]. However, in the coexistence phase (region P4), the stable fixed point can be either a node or a spiral point. In the second case, the transient dynamics of the system towards the fixed point may present some damped oscillations. The presence of such local oscillations is determined by the imaginary part of the eigenvalues of the model’s Jacobian matrix, calculated at the stable fixed point. For model B, which is two-dimensional, the Jacobian’s eigenvalues \( \sigma_1, \sigma_2 \) can either be both real or be complex conjugate to each other. For model A, which is three dimensional, the \( 3 \times 3 \) Jacobian matrix can either have none or have two nonreal conjugate eigenvalues. For both models A and B, one can use the quantity

\[
Q = \max_{i=1,\ldots,d} \left| \frac{\text{Im}(\sigma_i)}{\text{Re}(\sigma_i)} \right|, \tag{10}
\]

to measure the “quality factor” of the oscillations when \( \text{Re}(\sigma_i) \neq 0 \) (\( \sigma_i \) are the eigenvalues of the Jacobian at the fixed point, and \( d \) is the dimensionality of the system). This is because the imaginary part is responsible for the oscillations and the real part is responsible for the damping; thus the imaginary-to-real part ratio measures the propensity of the system to oscillate around the fixed point.

In Fig. 3, together with the regular phases of the model, we show in (color-coded) green the numerically calculated values of \( Q \). The coexistence phase can thus be subdivided according to the existence of a nonreal Jacobian eigenvalue: Within the white regions, the eigenvalues are real and the fixed point is a node, whereas in the green areas, the fixed point is a spiral point and there may occur oscillations around it. Notice
that for both models A and B, the shape of the spiral point region depends considerably on the time scale parameter $\pi$. Greater values of $\pi$ seem to shrink the oscillatory region and reduce the values of $Q$ for model A, but the opposite occurs for model B. Furthermore, we show the difference in Fig. 8, in which the time evolution of the prevalences of model B is shown for two situations: one with $Q = 0$ [hence no local oscillations; see Fig. 8(a)] and another with $Q = 3.65$ [there are damped oscillations before reaching the steady state; see Fig. 8(b)]. Interestingly, we only had to change the relative time scale parameter $\pi$ to switch between the two situations.

An important fact is that damped oscillations can only occur when the interaction between the diseases is asymmetrical, for both models A and B. For model A, we numerically check this by observing that the oscillatory region of the phase diagram shrinks and disappears as $\Gamma_1$ or $\Gamma_2$ leaves the region for which interactions are asymmetric. For model B, it can be shown that the Jacobian’s eigenvalue equation (which is quadratic) can only assume nonreal solutions if $\alpha > 1$. Finally, we note that despite that our deterministic formulation predicts that the oscillations are always damped, stochasticity—which is intrinsic to real-world systems and is inherent when performing Monte Carlo simulations—would cause such oscillations to last for a long time, as small perturbations to the prevalences could recover the oscillatory pattern.

### IV. CONCLUSIONS

In this paper, we have studied two minimalist models for interacting diseases in the asymmetrical regime, for homogeneously mixed populations and with continuous-time evolution. We focus on the influence of the relative time scale between the two diseases. The simplicity of our framework not only provides us with analytical tractability but also reveals the fundamental properties of asymmetrically interacting contagion. The models are simple enough to be applied to different situations, and the choice for two models is justified by our goal of achieving a more general understanding of interacting processes.

Model A is a mathematical prototype for an epidemic with awareness and the possible asymmetrical interaction between HIV and some specific diseases [29–31]. Our results show that despite the epidemic thresholds not being affected by the time scale parameter $\pi$, the prevalence is influenced in a nonintuitive way: The “prey” disease (disease I) has a greater prevalence when the “predator” disease (disease II) has a faster time scale. This behavior is in consonance with previous works on epidemics with awareness in complex networks [36,37], suggesting that it may be intrinsic to this class of models. If we take HIV, which naturally has a long time scale for its development, detection, and treatment, as an example, this means that the information awareness (which is more quickly transmitted and forgotten) may not be as efficient as it could be due to its shorter time scale.

Model B, which is inspired in situations such as computer viruses and spreading countermeasures [38–41] or fake vs fact-checked news, has a different behavior with the relative time scale: If the “predator” disease (disease II) has a faster clock, both prevalences of diseases I and II decay, with the possibility of disease I eradication [see Fig. 4(b)]. The interpretation then varies according to the situation being modeled. For example, if the main “goal” of disease II is to fight the other disease (such as fact-checking against fake news), then a shorter time scale (higher $\pi$) is desirable. If, however, the goal is to maximize disease II prevalence by taking advantage of disease I (one can imagine a computer virus that makes use of vulnerabilities generated by another one), then a longer time scale for disease II is preferable. This result is compatible with the analysis made by Wu et al. [44] using a heterogeneous mean field in a single-layer network.

Another interesting phenomenology revealed is the behavior of model B with its superinfection parameter $\alpha$, which controls the interaction between diseases I and II in both directions. The interpretation again depends on the situation: If the goal is to minimize disease I, then a greater value of $\alpha$ is always better, whereas for the optimization of the prevalence of disease II an intermediate value of $\alpha$ (which depends on $\pi$) should be sought. Finally, we have shown that oscillatory behavior, which is a well-known feature of predator-prey systems, is also present in asymmetrically interacting epidemic models and that its expression crucially depends on the relative time scale. Although the theory predicts that they are strongly damped, real-world systems could persistently display such oscillations due to random fluctuations.

To round off, we note that apart from the above connections to real dynamical systems, we expect that our work helps future studies of asymmetrically interacting spreading phenomena by providing general guidelines. By studying two different models, we provide a more general understanding of such systems, but we recognize that the results may have some bias due to the specific choice of models. An interesting extension of our work would be to consider a broader family of models, using more generic formulations, and determine, for example, the conditions that make the prevalences depend, in one way or another, on the time scale. Finally, we also expect some variations when these models are studied on top of structured populations. For such scenarios, our results provide...
a baseline of which qualitative and quantitative patterns are not associated with a network effect.

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APPENDIX A: ANALYTICAL EXPRESSIONS FOR THE COEXISTENCE FIXED POINT

The coexistence phase is characterized by the long-term presence of both diseases, thus having $u, v > 0$ in the steady state. In this Appendix, we show how to analytically calculate the prevalences for the coexistence fixed points of both models. For simplicity of notation, throughout Appendices A and B we use $u, v, w$ to denote the fixed point values of the prevalences (and not their time-dependent functions).

1. Model A

We find the fixed points by setting $\dot{u} = \dot{v} = \dot{w} = 0$ in the dynamical equations (1)–(3). Using the fact that $u, v > 0$ at the coexistence, we divide Eq. (1) by $(1 - \pi)\mu_1u$, Eq. (2) by $\pi\mu_2v$, and Eq. (3) by $(1 - \pi)\mu_1$, obtaining the following reduced system of equations:

\begin{align*}
0 &= \lambda_1[1 - u + (\Gamma_1 - 1)(v - w)] - 1, \quad (A1) \\
0 &= \lambda_2[1 - v + (\Gamma_2 - 1)(u - w)] - 1, \quad (A2) \\
0 &= \lambda_1\Gamma_1(v - w)u - w + \chi[\lambda_2\Gamma_2(u - w)v - w], \quad (A3)
\end{align*}

in which only Eq. (A3) is nonlinear. The disease II time scale factor $\chi$ is the same as defined in Eq. (10). Using also the definition

\[ s_i = 1 - \frac{1}{\lambda_i}, \quad i = 1, 2, \quad (A4) \]

where $s_i$ is the solution for a noninteracting SIS model with reproduction ratio $\lambda_i$, we can further simplify Eqs. (A1) and (A2) to

\begin{align*}
&u - (\Gamma_1 - 1)(v - w) = s_1, \quad (A5) \\
v - (\Gamma_2 - 1)(u - w) = s_2, \quad (A6)
\end{align*}

respectively, from which it is intuitive to see that the interaction between the diseases (represented by $\Gamma_1$ and $\Gamma_2$) causes a deviation from the noninteracting solution.

At this point, we split the solution into two cases: (a) the simpler case $\Gamma_1 = 0$ (i.e., when disease II completely inhibits infection by disease I) and (b) the more general case $\Gamma_1 > 0$.

\[ \text{a. Case } \Gamma_1 = 0 \]

In the case where $\Gamma_1 = 0$, Eq. (A5) simplifies to

\[ u - w = s_1 - v. \quad (A7) \]

Combining this equation with (A6), we get a $2 \times 2$ system for $v$ and $(u - w)$, for which the solution is

\[ v = \frac{s_2 + (\Gamma_2 - 1)s_1}{\Gamma_2}, \quad (A8) \]

\[ u - w = \frac{s_1 - s_2}{\Gamma_2}. \quad (A9) \]

Now using that $\Gamma_1 = 0$ in Eq. (A3), we obtain an expression for $w$ in terms of known variables:

\[ w = \frac{\chi}{1 + \chi} \lambda_2\Gamma_2(u - w)v. \quad (A10) \]

Plugging Eqs. (A8) and (A9) into the above equation and then using it back to Eq. (A9), we get the full expressions for the fixed points when $\Gamma_1 = 0$:

\begin{align*}
u &= s_1 - \left[1 - \frac{\chi}{1 + \chi} \frac{s_1 - s_2}{\Gamma_2} \left(s_2 + (\Gamma_2 - 1)s_1\right)\right], \quad (A11) \\
v &= \frac{s_2 + (\Gamma_2 - 1)s_1}{\Gamma_2}, \quad (A12) \\
w &= \frac{\chi}{1 + \chi} \frac{s_1 - s_2}{\Gamma_2} \left(s_2 + (\Gamma_2 - 1)s_1\right). \quad (A13)
\end{align*}

where we have also replaced $\lambda_2 = 1/(1 - s_2)$. With some analysis of the above expressions, noticing that $\chi/(1 - \chi)$ is an increasing function of $\chi > 0$, we can infer that $u$ and $w$ increase with $\chi$ (and thus with $\pi$), whereas $v$ does not depend on the time scale factor. This is in agreement with the plots in Figs. 5(a) and 5(b).

\[ \text{b. Case } \Gamma_1 > 0 \]

In the case where $\Gamma_1 > 0$, the fact that Eq. (A3) is quadratic on its variables cannot be avoided. Our strategy is to use Eqs. (A1) and (A2) to write $u, v, (u - w)$, and $(v - w)$ as functions of the variable $w$ and the model parameters. This can be used in Eq. (A3) to find the solution for $w$ and, consequently, for the other variables.

We can isolate $v$ in Eq. (A6) and apply it to Eq. (A5), obtaining

\[ u = m[P_{12} - (\Gamma_1 - 1)\Gamma_2w]. \quad (A14) \]

\[ v = m[P_{21} - (\Gamma_2 - 1)\Gamma_1w]. \quad (A15) \]

where we simplified the notation using the definitions

\begin{align*}
P_{12} &= s_1 + (\Gamma_1 - 1)s_2, \quad (A16) \\
P_{21} &= s_2 + (\Gamma_2 - 1)s_1, \quad (A17) \\
m &= 1/[1 - (\Gamma_1 - 1)(\Gamma_2 - 1)]. \quad (A18)
\end{align*}

We can also manipulate Eqs. (A14) and (A15) to obtain

\[ u - w = m[P_{12} - \Gamma_1w], \quad (A19) \]

\[ v - w = m[P_{21} - \Gamma_2w]. \quad (A20) \]
Now we plug the above expressions into Eq. (A3), which, after redistributing the terms and dividing them by $\Gamma_1 \Gamma_2 m^2$ ($m > 0$ for asymmetrical interactions and, by hypothesis, $\Gamma_1 > 0$), becomes the quadratic equation $aw^2 + bw + c = 0$, where

$$a = \lambda_1 (\Gamma_1 - 1) \Gamma_2^2 + \chi \lambda_2 (\Gamma_2 - 1) \Gamma_1, \quad (A21)$$

$$b = -\left\{ \lambda_1 [P_{12} + (\Gamma_1 - 1) P_{21}] + \chi \lambda_2 [P_{21} + (\Gamma_2 - 1) P_{12}] + \frac{1 + \chi}{m^2 \Gamma_1 \Gamma_2} \right\}, \quad (A22)$$

$$c = \frac{P_{12} P_{21}}{\Gamma_1 \Gamma_2} (\Gamma_1 \lambda_1 + \chi \Gamma_2 \lambda_2). \quad (A23)$$

For asymmetrical interactions, it is possible that $a = 0$ in the coexistence region. Thus we write the coexistence fixed point of the system as

$$w = \begin{cases} \frac{b - \sqrt{b^2 - 4ac}}{2a}, & a \neq 0 \\ a, & a = 0, \end{cases} \quad (A24)$$

$$u = m[P_{12} - (\Gamma_1 - 1) \Gamma_2^2 w], \quad (A25)$$

$$v = m[P_{21} - (\Gamma_2 - 1) \Gamma_1^2 w]. \quad (A26)$$

From the above expressions, it is difficult to extract the behavior of the prevalences with respect to the time scale parameter $\chi$ (or $\pi$). However, in Appendix B we present an alternative argument that reinforces the behavior observed in Fig. 5.

2. Model B

Model B has a much simpler procedure to find analytical expressions for the coexistence fixed point $u$ and $v$, for general values of the parameters. Knowing that $u, v > 0$, one can divide Eq. (4) by $(1 - \pi) \mu_1 u$ and Eq. (5) by $\pi \mu_2 v$ and set their left-hand sides to 0, obtaining

$$0 = \lambda_1 (1 - u - v) - 1 - \chi \alpha \lambda_2 v, \quad (A27)$$

$$0 = \lambda_2 (1 - u - v) - 1 + \alpha \lambda_2 u, \quad (A28)$$

which is a $2 \times 2$ linear system in $u$ and $v$. Passing convenient terms to the left side of each equation, one can write the system in terms of $s_1$ and $s_2$ as defined in (A4):

$$s_1 = u + \phi v, \quad (A29)$$

$$s_2 = (1 - \alpha) u + v, \quad (A30)$$

where we define $\phi$ as

$$\phi = 1 + \alpha \chi \lambda_2 / \lambda_1. \quad (A31)$$

The solution of this $2 \times 2$ system is

$$u = \frac{s_1 - \phi s_2}{1 + \phi (\alpha - 1)}, \quad (A32)$$

$$v = \frac{(\alpha - 1) s_1 + s_2}{1 + \phi (\alpha - 1)}. \quad (A33)$$

APPENDIX B: BEHAVIOR OF THE PREVALENCES

WITH $\pi$

By plotting the values of the fixed point prevalences, analytically derived in Appendix A, we could analyze the behavior of such prevalences with the time scale parameter $\pi$ for models A and B. In this Appendix, we provide analytical support for the observed behaviors in both models.

1. Model A

A possible approach to determine the slope of the fixed point prevalences $u, v, w$ with $\pi$ is to differentiate the expressions (A24)–(A26) with respect to $\chi$ [notice, from the definition in Eq. (10), that $\chi$ is an increasing function of $\pi$]. This procedure, however, can be very tedious and provides little or no analytical insight. An alternative approach is to implicitly differentiate the reduced equations (A1)–(A3) with respect to $\chi$, obtaining more insightful expressions.

Let us define here $u' = \partial u / \partial \chi$, $v' = \partial v / \partial \chi$, and $w' = \partial w / \partial \chi$. Implicit differentiation of Eqs. (A1)–(A3) yields

$$0 = -\lambda_1 u' + (\Gamma_1 - 1) \lambda_1 v' - (\Gamma_1 - 1) \lambda_1 w', \quad (B1)$$

$$0 = -\lambda_2 v' + (\Gamma_2 - 1) \lambda_2 u' - (\Gamma_2 - 1) \lambda_2 w', \quad (B2)$$

$$0 = \Gamma_1 \lambda_1 (v' - w') + \Gamma_2 \lambda_2 (v - w) u' - w' + [\Gamma_2 \lambda_2 (u - w) v' - w'] + \chi [\Gamma_2 \lambda_2 (u' - w') v + \Gamma_2 \lambda_2 (u - w) v' - w']. \quad (B3)$$

With some rearrangement, the above expressions can be written as a linear system given by

$$\mathbf{A} \mathbf{x}' = \mathbf{b}, \quad (B4)$$

where $\mathbf{x}' = (u', v', w')^T$ and

$$\mathbf{A} = \begin{bmatrix} -\lambda_1 & (\Gamma_1 - 1) \lambda_1 & - (\Gamma_1 - 1) \lambda_1 \\ (\Gamma_2 - 1) \lambda_2 & -\lambda_2 & - (\Gamma_2 - 1) \lambda_2 \\ A_{uu} & A_{uv} & A_{uw} \end{bmatrix}, \quad (B5)$$

with

$$A_{uu} = \Gamma_1 \lambda_1 (v' - w') + \chi \Gamma_2 \lambda_2 v', \quad (B6)$$

$$A_{uv} = \chi \Gamma_2 \lambda_2 (u - w) + \Gamma_1 \lambda_1 u, \quad (B7)$$

$$A_{uw} = \Gamma_1 \lambda_1 u + \chi \Gamma_2 \lambda_2 v + 1 + \chi. \quad (B8)$$

Moreover, the vector of independent coefficients is

$$\mathbf{b} = (0, 0, -(\Gamma_2 \lambda_2 (u - w) v' - w))^T = (0, 0, b_w)^T. \quad (B9)$$

Using Cramer’s rule, we can obtain the $\chi$ derivatives as functions of the model parameters and the prevalences:

$$u' = (b_w \lambda_1 \lambda_2 / \det \mathbf{A})(1 - \Gamma_1 \Gamma_2), \quad (B10)$$

$$v' = (b_w \lambda_1 \lambda_2 / \det \mathbf{A})(1 - \Gamma_2 \Gamma_1), \quad (B11)$$

$$w' = (b_w \lambda_1 \lambda_2 / \det \mathbf{A})(1 - (\Gamma_1 - 1)(\Gamma_2 - 1)). \quad (B12)$$

Thus, in the asymmetrically interacting regime ($0 < \Gamma_1 < 1$ and $\Gamma_2 > 1$), $u'$ has the same sign as $w'$ (thus $u$ and $w$ have
the same slope with \( \chi \) and \( \pi \), whereas \( \psi'(u) \) has opposite sign (slope). By showing that the ratio \( b_w/\det A \) is positive, we could demonstrate that \( u \) and \( w \) actually increase with \( \pi \), while \( v \) decreases. Although we could not mathematically determine the signals of \( b_w \) and \( \det A \) for arbitrary model parameters, we collected robust numerical evidence that \( b_w \) and \( \det A \) are both negative in the coexistence phase (region P4 in Fig. 3) for a wide set of model parameters, and so the ratio \( b_w/\det A \) is positive. This suggests that the behaviors shown in Figs. 4(a) and 5 are robust and should remain for the whole coexistence region.

2. Model B

For model B, one can extract the dependence of the prevalences with \( \pi \) (or \( \chi \)) directly from their analytical expressions [Eqs. (A32) and (A33)], noticing that \( \phi = 1 + \alpha \chi \lambda_2/\lambda_1 \) is an increasing function of \( \chi \). The prevalence \( v \) of disease II, as in Eq. (A33), is clearly a decreasing function of \( \phi \) for \( \alpha > 1 \) (which is the asymmetrically interacting case). From Eq. (A32), we can also directly infer that the prevalence \( u \) of disease I also decreases with \( \phi \) for \( s_2 > 0 \) (or equivalently, \( \lambda_2 \geq 1 \)). However, the coexistence phase (region P4) also comprehends a region at which \( \lambda_2 < 1 \), for which we should check the behavior with \( \phi \) more carefully. Taking the partial derivative of \( u \) with respect to \( \phi \), we get

\[
\frac{\partial u}{\partial \phi} = -s_2\left[1 + \phi(\alpha - 1) - (s_1 - \phi s_2)(\alpha - 1)\right]/[1 + \phi(\alpha - 1)^2].
\] (B13)

From the above expression, the condition \( \frac{\partial u}{\partial \phi} < 0 \) can be simplified as

\[
s_2 + (\alpha - 1)s_1 > 0,
\] (B14)

which is equivalent to the condition that \( \lambda_2 \) is above its critical value for coexistence, expressed by Eq. (8). This means that \( \frac{\partial u}{\partial \phi} < 0 \) in the whole coexistence region and, therefore, both prevalences \( v \) and \( u \) are decreasing functions of \( \phi, \chi, \) and \( \pi \). This is consistent with the observed behaviors in Figs. 4(b) and 6.

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