A contact process with mobile disorder

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Abstract. I study the absorbing-state phase transition in the one-dimensional contact process with mobile disorder. In this model the dilution sites, though permanently inactive, diffuse freely, exchanging positions with the other sites, which host a basic contact process. Even though the disorder variables are not quenched, the critical behavior is affected: the critical exponents $\delta$ and $z$, the ratio $\beta/\nu_\perp$ and the moment ratio $m = \langle \rho^2 \rangle / \rho^2$ take values different from those of directed percolation, and appear to vary with the vacancy diffusion rate. While the survival probability starting from a single active seed follows the usual scaling, $P(t) \sim t^{-\delta}$, at the critical point, the mean number of active sites and mean square spread grow more slowly than power laws. The critical creation rate increases with the vacancy density $v$ and diverges at a value $v_c < 1$. The scaling behavior at this point appears to be simpler than for smaller vacancy densities.

Keywords: critical exponents and amplitudes (theory), finite-size scaling, phase transitions into absorbing states (theory)
1. Introduction

Phase transitions to an absorbing state continue to attract much attention in nonequilibrium statistical physics [1]–[5], and have stimulated efforts to characterize the associated universality classes. Absorbing-state transitions arise in models of epidemics, population dynamics and autocatalytic chemical reactions, and underlie self-organized criticality in sandpile models. The simplest models exhibiting absorbing-state phase transitions are the contact process (CP) [1,6] and its discrete-time version, directed percolation (DP).

Scaling at absorbing-state phase transitions has been studied extensively, both theoretically and numerically, and more recently, in experiments [7]. A central conclusion is that critical behavior of the DP type is generic for models exhibiting a continuous phase transition to an absorbing state, in the absence of any additional symmetries or conserved quantities [8,9]. In particular, allowing particles to diffuse at a finite rate does not affect the critical behavior of the contact process [10].

The effect of quenched disorder on the CP, in the form of random dilution, as well as of randomly varying creation rates associated with each site, has been investigated in detail, beginning with the studies of Noest [11], and in subsequent analyses of the unusual scaling properties [12]–[20]. Little attention has been given to the effect of mobile disorder, for example in the form of diffusing passive sites. An exception is the recent study by Evron et al of a CP in which certain sites have a higher creation rate than others, with the ‘good’ and ‘bad’ sites diffusing on the lattice [21]. These authors report anomalies in the off-critical behavior of the model.

Lack of interest in the effect of mobile disorder at an absorbing-state phase transition presumably reflects a belief that it should not lead to any significant changes in scaling properties, and might in fact be irrelevant. In equilibrium, annealed disorder at fixed
A contact process with mobile disorder

concentration leads to ‘Fisher renormalization’ of the critical exponents if the specific heat exponent $\alpha > 0$ in the pure model, and does not alter the exponents if $\alpha < 0$ [22]. The condition for Fisher renormalization (FR), like that of Harris’s criterion [23], can be extended to models for which (as in the present case) $\alpha$ is not defined [24], by using the equivalent relation, $d\nu < 2$. (In the present context $\nu$ corresponds to the correlation length exponent $\nu_\perp$.) A subtlety nevertheless arises in extending the notion of FR to nonequilibrium systems, because this phenomenon requires equilibrium between the degrees of freedom associated with the disorder (such as vacancy positions) and the other elements comprising the system (e.g., spin variables). In the present case there is no energy function that could favor (via a Boltzmann factor $e^{-\beta H}$) one disorder configuration over another; all disorder configurations are equally probable, since the CP degrees of freedom have no influence over the vacancy dynamics. We therefore speak of ‘mobile’ rather than ‘annealed’ disorder. (Of course, those disorder configurations that are more favorable to the maintenance of activity in the CP will survive longer and so might dominate long-time averages over surviving realizations.)

In the FR scenario, the exponent $\alpha$ is replaced by $-\alpha/(1-\alpha)$, whilst the other static critical exponents (with the exception of $\eta$) are multiplied by a factor of $1/(1-\alpha)$. Annealed disorder is expected to be relevant for critical dynamics if $\nu(d+z) < 2$, where $z$ is the dynamic exponent [25]. In the case of the contact process, under the FR scenario, the critical exponents $\beta$ and $\nu_\perp$ would be renormalized by the same factor. (On the basis of the hyperscaling relation $\alpha = 2 - d\nu_\perp$, we would expect that $\beta \to \beta/(d\nu_\perp - 1)$, and similarly for the other static exponents.) It is less clear what would happen to the exponents governing critical dynamics, such as $z$ or $\delta$, or to an order parameter moment ratio such as $m \equiv \langle \rho^2 \rangle/\langle \rho \rangle^2$.

In this work I study the CP with mobile vacancies. The vacancies diffuse freely, at rate $D$, exchanging positions with nondiluted sites, be they active or inactive, which host the usual CP. Aside from the intrinsic interest, processes of this sort should find application in epidemic modeling (wandering of immune individuals in a susceptible population), and in population dynamics (for example, a diffusing population of a third species in a predator–prey system). The presence of diffusing impurities is in principle relevant to experimental realizations of directed percolation in fluid media [7]. It transpires that the critical creation rate $\lambda_c$ diverges at a certain vacancy density $v_c(D) < 1$, so an active state is impossible for $\nu > v_c$. It is thus of interest to study the limiting case of $\lambda \to \infty$, which corresponds to the limit of an epidemic that propagates extremely rapidly, and can be impeded only by fragmenting the population.

The remainder of this paper is organized as follows. In section 2 I define the model and discuss some of its basic properties as well as the results from the pair approximation. Section 3 presents the results from various simulation approaches, and in section 4 the implications for scaling behavior are discussed.

2. Model

In the basic contact process [6] each site of a lattice is in one of two states, ‘active’ or ‘inactive’. Transitions of an active site to the inactive state occur spontaneously (i.e., independent of the other states of the sites) at unit rate, while an inactive site becomes active at rate $\lambda n/q$, where $n$ is the number of active nearest neighbors and $q$ the total
number of neighbors. The configuration with all sites inactive is absorbing. The model is known to exhibit a continuous phase transition between the active and absorbing states at a critical creation rate $\lambda_c$; in the one-dimensional CP, $\lambda_c = 3.29785(2)$. (Figures in parentheses denote uncertainties.) The transition falls in the universality class of directed percolation. In the CP with mobile disorder (CPMV), a fraction $v$ of the sites are vacant, and do not participate in the usual CP dynamics. The vacancies hop on the lattice at rate $D$. In a hopping transition, a vacancy trades its position with its right or left neighbor. Denoting vacancies and the active and inactive sites by $v$, 1 and 0, respectively, we have the transitions $v1 \rightarrow 1v$ and $v0 \rightarrow 0v$ at rate $D/2$, and similarly for vacancy hopping to the left. (From here on, an ‘inactive’ site denotes one that is not a vacancy, but that happens not to be active.) The typical evolution shown in figure 1 illustrates that regions of higher than average vacancy density tend to be devoid of activity, and vice versa.

The CPMV is related to a model studied recently by Evron, Kessler and Shnerb (EKS) [21], in which half the sites are ‘good’, meaning that they have a creation rate $\lambda_1$, while the creation rate at the other, ‘bad’ sites is $\lambda_2 < \lambda_1$. The good and bad sites diffuse on the lattice at rate $D$, independently of their state (active or inactive). The CPMV resembles the limit $\lambda_2 \rightarrow 0$ of the EKS model, but the correspondence is not exact, since bad sites can in fact become active in the latter system, even if $\lambda_2 = 0$. (For $\lambda_2 = 0$, a bad site can never produce offspring, but it can always be activated by an active neighbor. In the CPMV by contrast, the vacancies can never be activated.)

While the EKS model with $\lambda_2 > 0$ represents a weaker form of disorder than the CPMV, one might expect the two models to fall in the same universality class. In principle,
the two models should have the same continuum limit, given by the following pair of stochastic partial differential equations:

\[ \frac{\partial \rho}{\partial t} = D \nabla^2 \rho + (a + \gamma \phi) \rho - b \rho^2 + \eta(x, t) \]  

(1)

and

\[ \frac{\partial \phi}{\partial t} = \bar{D} \nabla^2 \phi + \nabla \cdot \xi(x, t) \]  

(2)

where \( \rho(x, t) \) is the activity density and \( \phi(x, t) \) the density of nonvacant (or ‘good’) sites, whose evolution is completely independent of \( \rho \). Without the term \( \propto \phi \rho \), equation (1) is the usual continuum description of directed percolation [8, 9, 26], with \( \eta(x, t) \) a zero-mean, Gaussian noise with autocorrelation \( \langle \eta(x, t) \eta(y, s) \rangle \propto \rho(x, t) \delta(x - y) \delta(t - t') \). The second term on the rhs of equation (2) similarly represents a conserved delta-correlated noise.

It is not my intention to analyze these equations here. They merely serve to highlight the observation that if, as seems plausible, they represent the continuum limit for both models, then the models should have a common scaling behavior.

Suppose that, as seems reasonable, the CPMV exhibits a continuous phase transition to the absorbing state at a critical creation rate \( \lambda_c(v, D) \). For any nonzero vacancy concentration, \( \lambda_c \) must diverge as \( D \to 0 \), since we are then dealing with a CP on the line with frozen dilution, which has no active phase for finite \( \lambda \). The limiting value of \( \lambda_c \) as \( D \to \infty \) is a subtler issue. If all three species were to diffuse infinitely rapidly, the system would be well mixed, and would exhibit mean-field-like behavior. Note however that in one dimension, vacancy diffusion does not change the ordering of the active and inactive sites. Thus infinitely rapid vacancy diffusion does not correspond to a mean-field limit, but rather to a system in which a fraction \( v \) of sites are momentarily replaced, independently at each instant and at each lattice site, by inert elements. Hence we should expect to observe the behavior of the pure model, at an effective reproduction rate \( \lambda' = (1 - v) \lambda \), so the critical point is renormalized to \( \lambda'_c = \lambda_{c, \text{pure}}/(1 - v) \) in this limit.

While disorder that diffuses infinitely rapidly is irrelevant to critical behavior, one may argue that it is relevant for finite \( D \). Consider a correlated region in the CP, with characteristic size \( \xi \) and duration \( \tau \). If fluctuations in the vacancy density on this spatial scale relax on a timescale \( \tau_\phi \ll \tau \), then the CP will be subject, effectively, to a disorder that is uncorrelated in time, which is irrelevant [27]. But since the fluctuations relax via diffusion, \( \tau_\phi \sim \xi^2 \). In the neighborhood of the critical point, \( \xi \sim |\lambda - \lambda_c|^{-\nu_\perp} \) and \( \tau \sim \xi^z \), with \( z = \nu_\parallel / \nu_\perp \), so \( \tau_\phi \sim \tau^{2/z} \), suggesting that diffusing disorder is relevant for \( z < 2 \), provided of course that quenched disorder is relevant. (The two conditions are conveniently written as \( d\nu_\perp < 2 \) and \( d\nu_\parallel < 4 \).) In directed percolation this inequality is satisfied in \( d < 4 \) space dimensions.

It is perhaps worth mentioning that the present model is not equivalent to the diffusive epidemic process (DEP) [28]. In the latter model, there are two kinds of particles (healthy and infected, say), which diffuse on the lattice. Infected particles recover spontaneously while a healthy particle may become infected if it occupies the same site as an infected particle. (There is no limit on the occupation number at a given site in the DEP.) While it is tempting to identify the active and inactive sites of the CPMV with (respectively) the infected and healthy particles of the DEP, the analogy is not valid since, in the CPMV,
vacancy diffusion does not cause the nondiluted sites to change their relative positions, as noted above.

A simple and often qualitatively reliable approach to estimating the phase diagram is via \(n\)-site approximations, where \(n = 1\) corresponds to simple mean-field theory. At order \(n\), the \(m\)-site joint probability distributions (for \(m > n\)) are approximated on the basis of the \(n\)-site distribution [29, 30]. Since simple mean-field theory is insensitive to diffusion, the lowest order of interest here is \(n = 2\) (the pair approximation), which is derived in the appendix. The resulting transition line, \(\lambda_c(D)\), is compared with simulation results (for \(v = 0.1\)) in figure 2. The pair approximation turns out to be in rather poor agreement with simulation. In particular, it fails to show that \(\lambda_c\) diverges as \(D \to 0\). Some improvement might be expected using larger clusters, but it is not clear whether the correct asymptotic behavior would be captured even for large \(n\).

For a fixed diffusion rate \(D > 0\), we expect \(\lambda_c\) to increase monotonically with vacancy concentration \(v\). One may ask whether \(\lambda_c\) remains finite for any \(v < 1\), or whether it diverges at some vacancy concentration that is strictly less than unity. Now, for very large \(\lambda\), a string of two or more active sites is essentially immortal, for each time a site becomes inactive (which occurs at rate 1), it is immediately reactivated by its neighbor. Thus only isolated (nondiluted) sites can remain inactive. For \(v\) sufficiently large, we expect an isolated active site to become inactive before it can make contact with another nondiluted site, so \(\lambda_c \to \infty\) at some critical vacancy concentration \(v_c(D) < 1\). The following simple argument furnishes an estimate of \(v_c\). The rate of loss of activity due to isolated active sites becoming inactive is \(\simeq \rho v^2\) (that is, the mean-field estimate for the
density of isolated active sites). On the other hand, the rate of reactivation of inactive sites may be estimated as $D\rho v (1 - \rho - v)$, i.e., the vacancy diffusion rate times the mean-field estimate for the density of active–vacant–inactive triplets. The resulting equation of motion, $d\rho/dt = -v^2 \rho + D\rho v (1 - \rho - v)$, has the stationary solution $\rho = 1 - v - v/D$, so the activity density vanishes at $v = D/(1 + D) \equiv v_c$. While this approximation is too crude to yield a quantitative prediction, the expression for $v_c$ does at least attain the correct limiting values for $D \to 0$ and $D \to \infty$. (The pair approximation does not yield a useful prediction for $v_c$.) Simulation results for $v_c$ are reported in section 3.4.

3. Simulations

I performed three series of simulations of the model. The first series follows the evolution starting from a maximally occupied configuration (all nonvacant sites occupied) from the initial transient through the quasistationary (QS) state, in which mean properties of surviving realizations are time independent. These simulations will be referred to as ‘conventional’, to distinguish them from the second series, which employ the QS simulation method [33]. The third series studies the spread of activity starting from an initial seed. In all cases, the initial positions of the $vL$ vacancies are selected at random from the $L$ sites of the ring. A new selection is performed at each realization.

3.1. Conventional simulations

In conventional simulations (CS), the fraction $\rho$ of active sites and the moment ratio $m = \langle \rho^2 \rangle/\rho^2$ are calculated as functions of time. The time increment associated with a given event is $\Delta t = [(1 + \lambda) N_a + DN_v]^{-1}$, where $N_a$ and $N_v$ denote the numbers of active sites and of vacancies, respectively. $\rho(t)$ and $m(t)$ are calculated as averages over surviving realizations in time intervals that for large $t$ represent uniform intervals of $\ln t$, a procedure sometimes called ‘logarithmic binning’. In addition to the QS values of $\rho$ and $m$, conventional simulations furnish the mean lifetime $\tau$ from the asymptotic decay of the survival probability, $P_s \sim e^{-t/\tau}$. Estimates for the critical exponents $\delta$ and $z$ are obtained using the relations $\rho(t) \sim t^{-\delta}$ and $m(t) - 1 \sim t^{1/z}$ [31]. The latter relations describe the approach to the QS regime in a large system at its critical point. Finite-size scaling (FSS) theory implies that, at the critical point, $\rho_{QS} \sim L^{-\beta/\nu_1}$, while $\tau \sim L^z$, and that $m$ converges to a finite limiting value as $L \to \infty$ [32]. Note that we have, in principle, two independent ways to estimate the exponent $z$, one involving the scaling of the lifetime with system size, the other using the approach of $m$ to its QS value. (These scaling behaviors have been verified at absorbing-state transitions in various models without disorder; scaling in the presence of quenched disorder is generally more complicated. Whether simple FSS applies for mobile disorder is an open question; the results shown below provide partial confirmation.)

I performed CS for system sizes $L = 100, 200, \ldots, 1600$, averaging over $N_r$ independent realizations, where $N_r = 100000$ for $L = 100$, and decreases gradually with system size, to a value of 2000–5000 for $L = 1600$. The run times are such that all realizations eventually fall into the absorbing state. In analyzing the CS data, I use two criteria for determining the critical reproduction rate $\lambda_c$: (1) power-law scaling of the QS density, $\rho_{QS}$, with system size, and (2) approach of the moment ratio $m_{QS}$ to a well defined limiting value.
with increasing system size. Verification of (1) is facilitated by studying the curvature of $\ln \rho_{QS}$ versus $\ln L$, by calculating a quadratic fit to the data, and determining the range of $\lambda$ for which the quadratic term is zero to within uncertainty. It is also useful to plot $L^{\beta/\nu_{\perp}} \rho_{QS}$ versus $L$, on log scales, to check for curvature visually, as illustrated in figure 3. (Here $\beta/\nu_{\perp}$ is estimated from the data for the value of $\lambda$ showing the least curvature.) Verification of the second criterion is facilitated by plotting $m$ versus $1/L$; for subcritical values this graph veers upward as $1/L \to 0$ and vice versa, as shown in figure 4.

The behavior of the activity density as a function of time, at criticality, is shown in figure 5, for the case $v = 0.1$ and $D = 1$. In the pure contact process the activity follows a power law, with very weak corrections to scaling, for comparable times and system sizes [1]. Here by contrast, there is a significant positive curvature to the plots of $\rho(t)$ on log scales. It is nevertheless possible to estimate the exponent $\delta$ from the data at longer times and larger system sizes. Since the graphs of $\rho(t)$ are not linear (on log scales) it is not possible to collapse the data onto a unique master curve describing both the stationary and transient regimes, as is in fact possible in the contact process without dilution. The quantity $m(t) - 1$ shows cleaner power-law scaling over a somewhat larger period, as shown in figure 6, allowing a rather precise estimate for the exponent $z$.

Table 1 summarizes the results of the conventional simulations. The critical parameters such as $\beta/\nu_{\perp}$ and $m$ are calculated for each value of $\lambda$ simulated, and then evaluated at $\lambda_c$ using linear interpolation. Each fit to the data involves an error bar, to which we must add the (typically larger) uncertainty induced by the uncertainty in $\lambda_c$ itself. In table 1, the values cited for the dynamic exponent $z$ are obtained from the growth of the moment ratio (using $m - 1 \sim t^{1/z}$), and not from the FSS relation for the lifetime, $\tau \sim L^z$. In fact, the growth of the lifetime appears to be slower than a power law at the critical point, as shown in figure 7 for $v = 0.1$ and $D = 1$. In all the cases shown...
Figure 4. Moment ratio $m = \langle \rho^2 \rangle / \rho^2$ versus inverse system size in conventional simulations for $v = 0.1$, $D = 2$, and (upper to lower) $\lambda = 3.905$, 3.910, 3.915, and 3.920.

Figure 5. Order parameter $\rho(t)$ for $v = 0.1$, $D = 1$, and $\lambda = 4.099$. System sizes (upper to lower) $L = 200, 400, 800$, and 1600. The slope of the straight line is $-0.0863$.

(including $\lambda = 4.101$, above the critical value), the plots of $\tau^* = L^{-z} \tau$ curve downward. (In this plot, for purposes of visualization I use $z = 2.4$, derived from a linear fit to the data for $\tau(L)$ at $\lambda = 4.099$.)

The above results characterize the scaling behavior at the critical point. To estimate the critical exponent $\beta$, I study the supercritical regime. For $v = 0.1$ and $D = 0.5$, I
Figure 6. Evolution of $m(t) - 1$ for $v = 0.1$, $D = 1$, and $\lambda = 4.099$. System sizes as in figure 5. The slope of the straight line is 0.40.

Figure 7. Scaled lifetime $\tau^* = L^{-\beta} \tau$ versus system size for $v = 0.1$, $D = 1$, and (lower to upper) $\lambda = 4.097$, 4.099, and 4.101.

determined the order parameter $\rho$ for $\tilde{\Delta} \equiv (\lambda - \lambda_c)/\lambda_c = 0.002, 0.005, 0.01, 0.02, 0.05, 0.1, \text{ and } 0.2$, for systems of size 100, 200, ..., 3200. Using the relation $d \ln \rho/d\lambda|_{\lambda_c} \sim L^{1/\nu_\perp}$, with the derivative estimated via a linear fit to the data near the critical point ($\lambda_c = 4.375$), I obtain $1/\nu_\perp = 0.612(14)$, which combined with the result $\beta/\nu_\perp$ yields $\beta = 0.285(11)$. Plotting all of the data in the form of $\rho^* \equiv L^{\beta/\nu_\perp} \rho$ versus $\Delta^* = L^{1/\nu_\perp} \tilde{\Delta}$ leads to a reasonably good data collapse, as shown in figure 8. A fit to the data for $\ln \Delta^* \geq 2$ to

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A contact process with mobile disorder

Figure 8. Order parameter scaling plot, $\rho^* \equiv L^{\beta/\nu_{\perp}} \rho$ versus $\Delta^* = L^{1/\nu_{\perp}}[(\lambda - \lambda_c)/\lambda_c]$, for $v = 0.1$ and $D = 0.5$. System sizes $L = 100$ (open squares); 200 (filled squares); 400 (diamonds); 800 (open circles); 1600 (filled circles); 3200 (+).

Table 1. Critical parameters of the CPMV with vacancy density $v = 0.1$. $\delta_c$ and $\delta_s$ denote, respectively, the exponent as determined in conventional and in spreading simulations. DP values from [36].

| $D$  | $\lambda_c$ | $\beta/\nu_{\perp}$ | $m_c$  | $z$    | $\delta_c$ | $\delta_s$ |
|------|--------------|----------------------|--------|--------|-------------|------------|
| 0.5  | 4.375(2)     | 0.175(3)             | 1.076(2) | 2.65(4) | 0.0745(15) | 0.0765(2) |
| 1    | 4.099(1)     | 0.191(3)             | 1.085(2) | 2.49(1) | 0.086(2)   | 0.0837(5) |
| 2    | 3.915(1)     | 0.205(3)             | 1.096(3) | 2.36(5) | 0.101(4)   | 0.0971(5) |
| 5    | 3.7746(10)   | 0.235(4)             | 1.123(4) | 1.92(2) | 0.137(3)   | 0.1293(5) |
| DP   | 3.29785(2)   | 0.2521(1)            | 1.1736(1) | 1.58074(4) | 0.15947(3) | ($=\delta_c$) |

The form $\ln \rho^* = \beta \ln \Delta^* + \text{const}$ yields $\beta = 0.255(10)$, somewhat lower than the estimate obtained using the data near the critical point.

For $v = 0.1$ and $D = 1$, I performed studies that yield estimates of $\lim_{L \to \infty} \rho(\lambda, L) \equiv \rho(\lambda)$ in the supercritical regime. This is done using the observation that for large $L$, the order parameter approaches a limiting value via the relation $\rho(\lambda, L) = \rho(\lambda) + C/L$. I study $\lambda = 4.11–4.6$, corresponding to $\Delta = 0.0027–0.122$. (For $\lambda = 4.11$ data for systems of up to 25,600 sites are needed for a reliable extrapolation.) Using these data I obtain $\beta = 0.308(6)$. Analysis of the data for $d \ln \rho/d\lambda|_{\lambda_c}$ yields $1/\nu_{\perp} = 0.624(9)$, leading to $\beta/\nu_{\perp} = 0.192(6)$, in good agreement with the estimate obtained above using finite-size scaling, $\beta/\nu_{\perp} = 0.191(3)$.

Thus for $D = 0.5$ and $D = 0.5$ I find $\beta$ values that are not very different from the DP value of 0.2765. Note that the value under Fisher renormalization, assuming the validity of the hyperscaling relation $\alpha = 2 - d \nu_{\perp}$, would be very much larger: $\beta/(d \nu_{\perp} - 1) \simeq 2.86$ for DP in one spatial dimension.

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As noted in section 1, averages restricted to surviving realizations may yield statistical properties of the vacancies that differ from those of a random mixture. I studied the density of nearest-neighbor vacancy pairs; for a random mixture \( \rho_{vv} = N(N-1)/L^2 \), where \( N = vL \) is the number of vacancies. At the critical point (\( \lambda = 4.099, D = 1 \), and \( v = 0.1 \)), I find that \( \rho_{vv} \) grows with time, reaching a value about 2% greater than the random mixture value in a system with \( L = 100 \). For system sizes \( L = 200 \) and 800, the increase over the random mixture value is 1% and 0.2%, respectively. Thus selection effects on the vacancy configuration appear to be quite weak.

3.2. Diffusion rate zero

In the absence of vacancy diffusion, the system is divided into independent finite domains and there is no active stationary state. For \( \lambda > \lambda_{c, pure} \) and \( D = 0 \) the system is in a Griffiths phase in which the properties at long (but finite) times are dominated by exponentially rare large domains. The probability of a domain of size \( \ell \) decays \( \sim e^{-c\ell} \) while the lifetime of the CP grows as \( \exp[b\Delta_\perp \ell] \), leading to a power-law decay of the survival probability [11,12]. At very small diffusion rates one might expect to observe a crossover from Griffiths-phase scaling to that associated with the absorbing-state phase transition.

For \( D = 0.5 \) (the smallest diffusion rate studied here) and \( v = 0.1 \), the critical point falls at \( \lambda_c = 4.375 \). Figure 9 compares the survival probabilities \( P(t) \) (starting with all nonvacant sites active) with \( D = 0.5 \) and 0 for these values of \( \lambda \) and \( v \), in systems of 400 sites, in averages over 1000 realizations. The power-law decay of \( P(t) \) is evident in the absence of diffusion, while in the system with mobile vacancies the decay is exponential. (For these parameters, \( P(t) \sim t^{-0.47} \) for \( t \geq 10^4 \); reducing the creation rate to \( \lambda = 4 \), the decay exponent increases to 0.83.) Thus beyond a certain time, the survival probability is greater in the system without diffusion. The reason is that initially large domains remain large in this case, while for \( D > 0 \) the domain sizes fluctuate. The inset of figure 9 shows that the activity density \( \rho(t) \) (averaged over surviving realizations) is larger in the system with \( D > 0 \). The behavior of \( \rho(t) \) in the two cases is distinct for \( t \geq 20 \) or so. Recalling that the estimates for scaling properties are derived from the QS regime (i.e., where \( \rho(t) \) is essentially constant), it appears unlikely that these results are influenced by the presence of the Griffiths phase. (The influence of the latter might well be longer-lived at diffusion rates smaller than those studied here.) The surprising growth in \( \rho(t) \) at late times, in the system with nondiffusing vacancies, can be understood as a selection effect: as time increases, the set of surviving realizations becomes more restricted to those vacancy configurations having a large domain, on which the activity density is larger.

3.3. Quasistationary simulations

This simulation method is designed to sample directly the quasistationary (QS) regime of a system with an absorbing state, as discussed in detail in [33]. The method involves maintaining a list of active configurations during the evolution of a given realization. When a transition to the absorbing state is imminent, the system is instead placed in one of the active configurations, chosen randomly from the list. This procedure has been shown to reproduce faithfully the properties obtained (at much greater computational expense) in conventional simulations. The theoretical basis for the QS simulation method
remains valid in the presence of variables that continue to evolve even in the absence of activity, such as the vacancy positions in the CPMV, or the positions of healthy particles in the diffusive epidemic process [34].

I performed QS simulations on rings of $L$ sites, as in the conventional simulations. The simulation yields a histogram of the accumulated time during which the system has exactly $1, 2, \ldots, n, \ldots$, active sites, which is used to calculate the QS order parameter $\rho_{\text{QS}}$ and moment ratio $m_{\text{QS}}$. The lifetime $\tau$ is given by the mean time between attempted visits to the absorbing state, in the QS regime. Results are obtained from averages over 10 independent realizations, each lasting $t_{\text{max}} = 10^8$ time units, with an initial portion ($10^4-10^5$ time steps, depending on the system size), discarded to ensure the system has attained the QS state. One thousand saved configurations are used to implement the QS simulation procedure [33]. Values of the replacement probability $p_{\text{rep}}$ range from $10^{-3}$ to $6 \times 10^{-5}$ (smaller values for larger $L$); during the initial relaxation period, $p_{\text{rep}} = 0.01$.

Comparing the results of QS and conventional simulations near the critical point, I find no significant differences for $v = 0.1$ and $D = 0.5$, in systems of up to 1600 sites. For $D = 1$, the results for $\rho$ and $m$ again agree to within uncertainty, but the QS simulations yield slightly smaller values for the lifetime than conventional simulations. (For $L = 800$, the difference is about 5%.) The difference can be understood if we suppose that certain initial vacancy configurations are more favorable than others for the survival of activity. In the conventional simulations each initial configuration makes a single contribution to the average survival time, whereas in QS simulations, realizations starting with less favorable configurations (with shorter lifetimes) visit the absorbing state more frequently than those starting from favorable ones, and so tend to dominate the average for the lifetime. This presumably reflects the slow relaxation of the vacancy density.
fluctuations. An advantage of conventional simulations in this respect is that many more initial configurations are sampled than in the QS simulations. (The former also provide information on the relaxation process, that is not available in QS simulations.) Despite the minor difference in lifetime estimates, the otherwise excellent agreement between the two simulation methods lends confidence to the results obtained.

3.4. Spreading simulations

In the light of the surprising results of the conventional simulations, it is of interest to study the model using a complementary approach, which follows the spread of activity, starting from a localized seed. Such spreading or ‘time-dependent’ simulations were pioneered in the context of contact process by Grassberger and de la Torre [35]. Starting with a single active site at the origin, one determines (as an average over a set of many realizations) the survival probability \( P(t) \), mean number of active sites \( n(t) \), and mean square spread, \( R^2(t) = \langle \sum_j x_j(t)^2 \rangle / n(t) \). (Here the sum is over all active sites, with \( x_j \) the position of the \( j \)th active site, and the brackets denote an average over all realizations.) The spreading process is followed up to a certain maximum time \( t_m \), with the system taken large enough that activity does not reach the boundaries up to this time. The expected scaling behaviors at the critical point are conventionally denoted as \( P(t) \sim t^{-\delta} \), \( n(t) \sim t^\eta \) and \( R^2(t) \sim t^{z_2} \).

Away from the critical point, these quantities show deviations from power laws. \( P(t) \), for example, decays exponentially for \( \lambda < \lambda_c \), and approaches a nonzero asymptotic value, \( P_\infty \), for \( \lambda > \lambda_c \).

In the CP with mobile disorder, this scaling picture appears to hold for the survival probability but not for \( n(t) \) and \( R^2(t) \). For example, figure 10 shows the decay of the survival probability for \( v = 0.1 \) and \( D = 2 \); \( P(t) \) is well fitted by an asymptotic power law with \( \delta = 0.0971(5) \). (Note however that at shorter times \( t < 100 \) the decay is governed by a larger exponent, with a value of about 0.147.) The small differences between estimates for \( \delta \) characterizing the decay of the order parameter in CS, and that of the survival probability in spreading simulations (\( \delta_c \) and \( \delta_s \), respectively, in table 1), can be attributed to the deviations from a pure power law noted above in the CS. (Note also that the uncertainties reported for \( \delta_s \) do not include any contribution due to the uncertainty in \( \lambda_c \), and so are smaller than the uncertainties in \( \delta_c \).)

The functions \( n \) and \( R^2 \) at the critical point are illustrated in figure 11, from which it is clear that their growth is not described by pure power laws. The mean square spread \( R^2 \) appears to cross over from a power law at short times to one with a smaller growth exponent at longer times. In fact an excellent fit to the data is obtained using the crossover form

\[ R^2 \propto t^{z_2} \frac{1}{1 + (t_x/t)^\phi} \tag{3} \]

where \( z_2 \) is the exponent at long times, \( \phi = z_1 - z_2 \) with \( z_1 \) the growth exponent at short times, and \( t_x \) is a crossover time. The best-fit parameters are \( z_1 = 1.25 \), \( z_2 = 0.77 \), and \( t_x = 6875 \). In the case of the mean number of active sites, \( n(t) \), I obtain an excellent fit with a long-time growth exponent of zero, and so use the simpler form

\[ n \propto \frac{1}{1 + (t_x/t)^\eta} \tag{4} \]
Figure 10. Lower curve: survival probability $P$ versus time for $v = 0.1$ and $D = 2, \lambda = 3.915$, average over $10^4$ realizations. Upper curve: scaled survival probability $t^\delta P(t)$ using $\delta = 0.0971$.

Here the best-fit parameters are $\eta = 0.367$ and $t_{\times} = 1.05 \times 10^6$. Thus it is possible that at long times $n$ grows more slowly than a power law, or saturates at a finite value as $t \to \infty$. Simulations extending to longer times would be required to resolve this question, which I defer to future study. It is worth noting that the exponents governing the short-time evolution are not far from the associated DP values of $\eta = 0.3137$ and $z_s = 1.2652$. The exponent values found at long times do not, however, seem characteristic of any known universality class.

3.5. Critical vacancy concentration

At a fixed diffusion rate, one expects $\lambda_c$ to be an increasing function of $v$, as noted in section 2; simulation results for $D = 1$ confirm this. (For $v \geq 0.2$ I determine $\lambda_c$ in conventional simulations, using the criteria discussed above, in studies of systems with $L \leq 800$.) To determine the critical vacancy concentration $v_c(D)$, at which $\lambda_c$ diverges, I perform conventional simulations at an effectively infinite creation rate. This is done by (1) allowing only isolated active sites to become inactive, at a rate of unity, and (2) activating any nondiluted site the instant it gains an active neighbor. (Thus a string of inactive sites are activated simultaneously when the rightmost or leftmost site acquires an active neighbor via vacancy motion.) These studies yield $v_c(1) = 0.517(1)$ for $D = 1$, and $v_c \simeq 0.4182(5)$ for $D = 0.2$. Figure 12 shows $\lambda_c(v)$ for $D = 1$; the data suggest that $\lambda_c \sim (v_c - v)^{-1.16}$ (see the inset). The available results are consistent with $\beta/\nu_\perp$, and the critical moment ratio $m_c$, being constant along the critical line for fixed $D = 1$.

I turn now to the scaling behavior at the critical vacancy density $v = v_c, \lambda = \infty$. For $D = 1$, finite-size scaling analysis of the order parameter yields $\beta/\nu_\perp = 0.184(20)$. Unlike in the case $v = 0.1$ discussed above, here the survival time follows a power law...
A contact process with mobile disorder

Figure 11. Mean number $n$ of active sites (lower curve) and mean square distance from origin $R^2$ (upper curve) versus time; parameters as in figure 10. Points: simulation; solid lines: fitting function (see the text).

Table 2. Critical parameters of the CPMV at the critical vacancy density $v_c$. $\delta_c$ and $\delta_s$ as in table 1. $z_m$ is the dynamic exponent as determined from the growth of $m$ in conventional simulations. DP values from [36].

| $D$ | $v_c$  | $\beta/\nu$  | $m_c$  | $z_m$  | $\delta_c$ | $\delta_s$ | $\eta$  | $z_s$  |
|-----|--------|--------------|--------|--------|------------|------------|--------|--------|
| 0.2 | 0.4182(5) | 0.174(6)     | 1.08(3) | 1.95(4) | 0.08(2)    | 0.086(2)   | 0.30(3) | 0.95(1) |
| 1   | 0.517(1)  | 0.184(20)    | 1.08(4) | 1.98(3) | 0.09(4)    | 0.086(2)   | 0.307(1) | 0.965(10) |
| DP  | —      | 0.2521(1)    | 1.1736(1) | 1.58(4) | 0.159(4)   | 0.313(4)   | 1.265(23) |  |

with $z = 1.98(3)$, close to the value for $z$ obtained via analysis of the growth of $m(t)$. The short-time scaling behavior $\rho \sim t^{-\delta}$ is followed quite clearly (see figure 13), without the strong crossover effects observed for $v = 0.1$ (figure 5). Figure 14 is a scaling plot of $m - 1$, using $t^* = t/L^{z_m}$, showing a near-perfect data collapse for three system sizes, and a clean power law, $m - 1 \sim t^{1/z}$. There is also good evidence for power-law scaling of the survival probability and of $n$ and $R^2$ at $v_c$, as shown in figure 15. Thus the scaling behavior at the critical vacancy density appears to be simpler than that observed for $v = 0.1$. Results for critical parameters at the critical vacancy density are summarized in table 2. For the two diffusion rates studied, the exponents and $m_c$ agree to within uncertainty. The spreading exponents satisfy the hyperscaling relation [35] $4\delta + 2\eta = dz$ to within uncertainty. The principal inconsistency in these data is in regard to the scaling relation $z = 2/z_s$. In fact $2/z_s$ is about 8% (5%) greater than $z_m$ for $D = 0.2$ ($D = 1$). Studies of larger systems and a more precise determination of $v_c$ should help to resolve this discrepancy.

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Figure 12. Critical creation rate $\lambda_c$ versus vacancy concentration $v$ for $D = 1$. The dashed vertical line marks $v_c$. Inset: the same data plotted versus $v_c - v$; the slope of the straight line is $-1.16$.

Figure 13. Order parameter versus time for $v = 0.5176$, $\lambda = \infty$, with $D = 1$. System sizes $L = 796$ (upper) and 1592 (lower). The slope of the straight line is $-0.093$. 

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Figure 14. Scaling plot of $m - 1$ versus $t^* = t/L^z$ using $z = 1.98$. Parameters $v = 0.5176$, $\lambda = \infty$, and $D = 1$. System sizes $L = 398$ (+), $L = 796$ ($\times$), and 1592 (circles). The slope of the straight line is 0.51.

Figure 15. Spreading simulations at the critical vacancy density. Solid lines (lower to upper): $P(t)$, $n(t)$ and $R^2(t)$. The corresponding dotted lines show the scaled quantities $P^* = t^{\delta} P$, $n^* = t^{-\eta} n$, and $R^{*2} = t^{-z_s} R^2$, using $\delta = 0.086$, $\eta = 0.307$, and $z_s = 0.965$. Parameters $v = 0.516$, $\lambda = \infty$, $D = 1$, system size $L = 6000$. 

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4. Discussion

I study the one-dimensional contact process with mobile vacancies, verifying that the model exhibits a continuous phase transition between the active and the absorbing phases, as in the undiluted model. The scaling behavior is, however, considerably more complex than the Fisher renormalization scenario that might be expected by analogy with equilibrium systems with annealed disorder. (The analogy, as noted, is somewhat flawed since the disorder variables are not in equilibrium with the other degrees of freedom.) Cluster approximations appear to be incapable of providing an accurate phase diagram. Quantitative results on critical behavior are obtained using conventional simulations (CS) and spreading studies. Quasistationary simulations yield results consistent with CS, except for a slightly smaller survival time.

The scaling behavior is anomalous in several aspects. First, the growth of the lifetime $\tau$ at the critical point appears to be slower than a power law, as is also the case for mean number of active sites and mean square spread in spreading simulations at the critical point. The latter finding is reminiscent of the one-dimensional CP with quenched disorder in the form of an annihilation rate randomly taking one of two values, independently at each site [14, 15]. In this case an intermediate phase arises, in which the survival probability, starting from a single active site, attains a nonzero value as $t \to \infty$, but the active region grows in a sublinear manner with time. The spreading simulations also show a crossover from scaling with exponent values similar to those of DP at short times, to a different behavior at longer times.

In section 2 it was argued that $\lambda_c \to \infty$ as $D \to 0$, and that $\lambda_c \to \lambda_{c,\text{pure}}/(1-v)$ in the opposite limit. The available data are consistent with this (see figure 2); extrapolation of the data for $v = 0.1$ yields a limiting ($D \to \infty$) value of $\lambda_c = 3.65$, quite near $\lambda_{c,\text{pure}}/(1-v) \simeq 3.664$.

For vacancy density $v = 0.1$, the critical exponents $\delta$ and $z$, the ratio $\beta/\nu_\perp$, and moment ratio $m_c$ seem to vary systematically with the vacancy diffusion rate (table 1). With increasing $D$ these parameters tend to their directed percolation values, as might be expected, since the limit $D \to \infty$ corresponds to the usual contact process, with a renormalized creation rate. This raises the question of whether the exponent values reported in table 1 reflect the true scaling behavior, or appear to vary as a result of corrections to scaling. In this regard it is pertinent to recall the scaling relation $\delta = \beta/\nu_\parallel = \beta/(\nu_\perp z)$. For the exponents in table 1, $\delta_c - \beta/(\nu_\perp z)$ deviates significantly from zero, taking values of 0.008(4), 0.009(4), 0.014(7) and 0.015(6) for $D = 0.5, 1, 2,$ and 5, respectively. The violation of scaling is smaller for the smaller $D$ values, suggesting that the associated exponent estimates are closer to the true values. Thus it is entirely possible that the apparent variation of exponents with $D$ reflects a correction to scaling, due perhaps to the expected crossover to DP scaling in the $D \to \infty$ limit. Resolution of this issue will require study of larger systems.

There is preliminary evidence that some critical properties are insensitive to varying the vacancy concentration $v$ at a fixed diffusion rate. Preliminary results for critical behavior at the critical vacancy density $v_c$ suggest a simpler scaling behavior, without strong crossover effects, at this point. Indeed, the scaling relation $\delta = \beta/(\nu_\perp z)$ is satisfied to within uncertainty for the exponent estimates in table 2. The static, transient, and
spreading behavior is quite analogous to that observed in the basic contact process, but with a different set of critical exponents. This motivates the conjecture that for \( v > 0 \) all points along the critical line \( \lambda_c(v) \) flow, in the renormalization group sense, to the same fixed point. For smaller values of \( v \) one would then expect to observe a crossover between DP scaling and that of the new fixed point, associated with mobile disorder. More extensive studies are needed to verify and fill out this picture.

The effects of mobile reported here are more extensive than those found by Evron et al in a related model (EKS) [21]. These authors do not find, for example, changes in the static critical exponents. In comparing the two models, it should be noted that the disorder in the CPMV is stronger than that involved in the EKS model. The difference nevertheless raises an interesting puzzle regarding the universality of disorder effects.

Intuitively, the anomalous scaling properties of the CP with mobile disorder are associated with fluctuations in the local, time-dependent concentration of vacancies, which relax slowly in the long-wavelength limit. Developing a quantitative theory of how such fluctuations change the critical behavior is a challenge for future work, with potential applications in epidemiology and the dynamics of spatially distributed populations. Investigation of the two-dimensional case is also relevant for applications, and of intrinsic interest for developing a more complete understanding of scaling.

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Appendix. Pair approximation

It is straightforward to construct the pair approximation to the CPMV. For convenience I denote vacant, inactive, and active sites by \( v, 0, \) and 1, respectively. In this approximation the dynamical variables are the probabilities \( p(00) \equiv \langle 00 \rangle, \langle 01 \rangle, \) etc. Note that since vacancies hop independently of the states (0 or 1) of the nonvacant sites, \( \langle vv \rangle = v^2 \) in the stationary state, and during the entire evolution, if this equality holds initially. Since \( v = \langle vv \rangle + \langle 0v \rangle + \langle 1v \rangle, \) we then have \( \langle 1v \rangle = v - v^2 - \langle 0v \rangle, \) leaving four independent pair probabilities. (The normalization condition, \( \langle 00 \rangle + \langle 11 \rangle + \langle vv \rangle + 2[\langle 01 \rangle + \langle 0v \rangle + \langle 1v \rangle], \) permits us to eliminate one further variable.)

Consider, for example, the transition \( 00 \to 01. \) It occurs at rate \( \lambda/2 \) provided the site to the right of the pair bears a particle, so the contribution to \( d\langle 00 \rangle/dt \) due to this event is \( -\langle \lambda/2 \rangle \langle 00 \rangle \langle 0v \rangle \). (The mirror event makes an identical contribution.) In the pair approximation, three-site probabilities are written in terms of the one- and two-site probabilities, for example, \( \langle 001 \rangle \approx \langle 00 \rangle \langle 01 \rangle / \langle 0 \rangle. \) Proceeding in this manner we obtain

\[
\frac{d\langle 00 \rangle}{dt} = 2\langle 01 \rangle + D \frac{(0v)^2}{v} - \frac{\langle 00 \rangle}{\langle 0 \rangle} \left[ \lambda \langle 01 \rangle + D\langle 0v \rangle \right] \\
\frac{d\langle 01 \rangle}{dt} = \langle 11 \rangle + \frac{\lambda}{2} \frac{\langle 00 \rangle}{\langle 0 \rangle} + D \frac{(0v)(1v)}{v} - \langle 01 \rangle \left[ 1 + \frac{D}{2} \left( \frac{(1v)}{(1)} + \frac{(0v)}{(0)} \right) + \frac{\lambda}{2} \left( 1 + \frac{(10)}{(0)} \right) \right]
\]

\( (A.1) \) (A.2)
A contact process with mobile disorder

\[
\frac{d(0v)}{dt} = (1v) + \frac{D}{2} \left[ \frac{(0)(0v)}{(0)} + v(0v) + \frac{(01)(1v)}{(1)} \right] - \frac{D}{2} \frac{(0v)}{v} \left[ \frac{(0v)}{v} + \frac{(1v)}{v} + \frac{(0v)}{(0)} \right] - \frac{\lambda}{2} \frac{(0v)}{(0)} \tag{A.3}
\]

and

\[
\frac{d(11)}{dt} = \lambda (01) \left( 1 + \frac{(10)}{(0)} \right) + D \frac{(1v)^2}{v} - (11) \left( 2 + D \frac{(1v)}{(1)} \right). \tag{A.4}
\]

Integrating these equations numerically, one can determine the critical creation rate \( \lambda_c \) as a function of \( v \) and \( D \).

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