Abstract
In tracking of time-varying low-rank models of time-varying matrices, we present a method robust to both uniformly-distributed measurement noise and arbitrarily-distributed “sparse” noise. In theory, we bound the tracking error. In practice, our use of randomised coordinate descent is scalable and allows for encouraging results on changedetection.net, a benchmark.

Introduction
Dimension reduction is a staple of Statistics and Machine Learning. In principal component analysis, its undergraduate-textbook version, possibly correlated observations are transformed to a combination of linearly uncorrelated variables, called principal components. Often, a low number of principal components suffice for the so-called low-rank model to represent the phenomenon observed. Notoriously, however, a small amount of noise can change the principal components considerably. A considerable effort has focussed on the development of robust approaches to principal component analysis (RPCA). Two challenges remained: robustness to both sparse and non-sparse noise and theoretical guarantees in the time-varying setting.

We present the pursuit of time-varying low-rank models of time-varying matrices, which is robust to both dense uniformly-distributed measurement noise and sparse arbitrarily-distributed noise. Consider, for example, background subtraction problem in Computer Vision, where one wishes to distinguish fast-moving foreground objects from slowly-varying background in video data (Liu et al. 2013). There, a matrix represents a constant number of frames of the video data, flattened to one row-vector per frame. At any point in time, the low-rank model is captured by a short-and-wide matrix. The time-varying low-rank model makes it possible to capture slower changes, e.g., lighting conditions slowly changing with the cloud cover. There may also be slight but rapid changes, e.g., leaves of grass moving in the wind, which could be captured by the uniformly-distributed dense noise. Finally, the moving objects are captured by the sparse noise. Clearly, low-rank modelling has wide-ranging applications beyond Computer Vision, wherever one needs to analyse high-dimensional streamed data and flag abnormal observations to operators, while adapting the model of what is normal over time.

Our contributions are as follows:

• we extend the low-rank + sparse model to low-rank + dense uniformly-distributed noise + sparse, where low-rank can be time-varying
• we provide an algorithm with convergence-rate guarantees for the time-invariant case
• we provide an algorithm with guarantees for the time-varying case. In Theorem 2, we bound the tracking error of an algorithm for any low-rank factorisation problem for the first time. That is: we show that a sequence of approximately optimal costs eventually reaches the optimal cost trajectory.
• we improve upon the statistical performance of RPCA approaches on changedetection.net of (Goyette et al. 2012), a well-known benchmark: the F1 score across 6 categories of changedetection.net improves by 28%, from 0.44643 to 0.57099. On the baseline category, it is 0.80254.
• we improve upon run time per frame of the same RPCA approaches, as detailed in Table 1. Compared to TTD_3WD, to give an example of a method which is still considered efficient in the literature, our single-threaded implementation is 103 times faster.

Related Work
Traditional approaches to robustness in low-rank models (Candès and Recht 2009) to name some of the pioneering work) are based on a long history of work in robust statistics (Huber 1981). In such approaches (Candès et al. 2011; Feng et al. 2013; Guo, Qiu, and Vaswani 2014; Mardani, Mateos, and Giannakis 2013), sometimes known as “Low-rank + Sparse”, one balances the number of samples of the “sparse” noise and the rank of the model, or the nuclear norm as a proxy for the rank. There are a number of excellent implementations, including some focused on the incremental update (Lin, Liu, and Su 2011; He, Balzano, and Lui 2011; Balzano and Wright 2013; Oreifej, Li, and Shah 2013; Meng and Torre 2013; Rodriguez and Wohlberg 2013; Dutta and Li 2017; Dutta, Li, and Richtarik 2017; Ma and Aybat 2018; Lerman and Maunu 2018; Vaswani and Narayana-
murthy 2018] Balzano, Chi, and Lu 2018] Yong et al. 2018, e.g.). In our comparison, we focus five of the best-known implementations and one very recent one. LRR_FastLADMAP [Lin, Liu, and Su 2011], RPCA_FPCP [Rodriguez and Wohlberg 2013], and MC_GROUSE (Balzano and Wright 2013) use the low-rank + sparse model. ST_GRASIA (He, Balzano, and Liu 2011) uses rank-1 + sparse. TTD_3WD (Orefeij, Li, and Shah 2013) uses low-rank + turbulence + sparse. The most recent formulation we consider is OMoGMF (Yong et al. 2018), which utilises a Gaussian mixture model (GMM) structure over the low-rank model, plus sparse noise on top. We refer to [Bhojanapalli, Neyshabur, and Srebro 2016 Boumal, Veroninski, and Bandeira 2016] Jain and Kar 2017 Boumal, Absil, and Curtis 2018] Bhojanapalli et al. 2018 for the present-best theoretical analyses in the off-line, time-invariant case, but stress that no guarantees have been known for the on-line, time-varying case. We refer to [Bhojanapalli, Neyshabur, and Srebro 2016; Bhojanapalli, Neyshabur, and Srebro 2016; Jain and Kar 2017; Jain and Kar 2017; Jain and Kar 2017; Jain and Kar 2017] for up-to-date surveys.

### Problem Formulation

Consider \( N \) streams with \( n \)-dimensional measurements, coming from \( N \) sensors with uniform sampling period \( h \) from \( t_k \) till \( t_k + hT \) (possibly with many missing values), packaged in a (possibly partial) matrix \( \mathbf{M}_k \in \mathbb{R}^{T \times nN} \). Every time a new observation comes in, its flattening is added at the bottom row to the matrix and the first row is discarded. In this way, the observation matrix slowly varies over time, i.e., \( \mathbf{M}_{k+1} \) is different from \( \mathbf{M}_k \), in general.

It is natural to assume that any row \( d \) may resemble a linear combination of \( r \ll T \) prototypical rows. Prior to the corruption by sparse noise, we assume that there exists \( \mathbf{R}_k \in \mathbb{R}^{r \times nN} \), such that flattened observations \( \mathbf{x}_d \in \mathbb{R}^{1 \times nN} \) are

\[
\mathbf{x}_d = \mathbf{c}_d \mathbf{R}_k + \mathbf{e}_d, \tag{1}
\]

where the row vector \( \mathbf{c}_d \in \mathbb{R}^{1 \times r} \) weighs the rows of matrix \( \mathbf{R}_k \), while \( \mathbf{e}_d \in \mathbb{R}^{1 \times nN} \) is the noise row vector, where each entry be uniformly distributed between known, fixed \(-\Delta\) and \(\Delta\). Further, this formulation (1) is extended towards the contamination model (Huber 1981), where “sparse errors” replace readings of some of the sensors. That is: Either we receive a measurement belonging to our model, or not:

\[
(x_d)_i = (I_n - \mathbb{1}_{i,k}) \circ [(c_d \mathbf{R}_k)_i + (e_d)_i] + \mathbb{1}_{i,k} \circ s_i, \tag{2}
\]

where index \( i \) enumerates sensors, \( s_i \in \mathbb{R}^{1 \times n} \) is a generic noise vector, while the Boolean vector \( \mathbb{1}_{i,k} \in \{0, 1\}^n \) has entries that are all zeros or ones depending on whether we receive a measurement belonging to our model or not. The operation \( \circ \) represents element-wise multiplication.

Considering the matrix representation, we assume that the matrix \( \mathbf{M}_k \) can be decomposed into slowly varying low-rank model \( (\mathbf{C}_k, \mathbf{R}_k) \) and additive deviation \( (\mathbf{E}_k) \) from the model comprising noise and anomalies:

\[
\mathbf{M}_k = \begin{bmatrix} \ldots & \mathbf{x}_d & \ldots \end{bmatrix} = \mathbf{C}_k \mathbf{R}_k + \mathbf{E}_k, \tag{3}
\]

where \( T \) is the number of samples stacked in rows of matrix \( \mathbf{M}_k \), \( r \) is the number of prototypes in the low-rank approximation, \( \mathbf{x}_d \) is a \( d \)-th row-vector in matrix \( \mathbf{M}_k \), \( \mathbf{C}_k \in \mathbb{R}^{T \times r} \) and \( \mathbf{E}_k \in \mathbb{R}^{T \times nN} \) are the matrices incorporating the coefficient vectors \( \mathbf{c}_d \)'s and noise \( \mathbf{e}_d \)'s as \( \mathbf{C}_k = \begin{bmatrix} \ldots ; \mathbf{c}_d ; \ldots \end{bmatrix} \), and

\[
\mathbf{E}_k = \begin{bmatrix} \ldots ; \mathbf{e}_d ; \ldots \end{bmatrix},
\]

respectively.

The missing entries in \( \mathbf{M}_k \) can represent either really absent data or outliers, such as moving objects in the case of video-processing applications. One can assume that normal behaviour exhibits certain regularity, which could be captured by a low-rank structure, and that events or anomalies are sparse across both time and space. The sparsity should be construed quite loosely, for example, comprising dense blobs of pixels moving coherently in video data, while occupying a relatively small fraction of image pixels in total. This notion of anomaly detection is widely used in monitoring streamed data, event recognition, and computer vision.

If we can identify the low-rank model, any deviation from the measurement model can be interpreted as an anomaly or event. When there are few measurements for which \( \mathbb{1}_{i,k} = 1_n \) and those are different from standard measurements, i.e., the aggregated \( \mathbb{1}_k \in \{0, 1\}^n \), which stacks all the individual \( \mathbb{1}_{i,k} \) for a specific time \( k \), is sparse, and samples of \( s_i \) fall outside of some range \([M_{k,ij}, M_{k,ij}]\) (defined below), it is possible to identify samples of \( s_i \) perfectly. In this paper, we provide a way to detect such anomalies, i.e., measurements for which \( \mathbb{1}_{i,k} = 1_n \). Hence, we are effectively proposing a
principal component pursuit algorithm robust to uniform and sparse noise.

We compute matrices $C_k$ and $R_k$ by resorting to a low-rank approximation of the matrix $M_k$ with an explicit consideration of the uniformly-distributed error in the measurements. Let $M_{k,i,j}$ be the $(i,j)$ element of $M_k$. Consider the interval uncertainty set $[M_{k,i,j} - \Delta, M_{k,i,j} + \Delta]$ around each observation. Finding $(C_k, R_k)$ can be seen as matrix completion with element-wise lower bounds $M_{k,i,j} := M_{k,i,j} - \Delta$ and element-wise upper bounds $M_{k,i,j} := M_{k,i,j} + \Delta$. Let $C_{k,i}$ and $R_{k,j}$ be the $i$-th row and $j$-th column of $C_k$ and $R_k$, respectively. With Frobenius-norm regularisation, the completion problem we solve is:

$$\min_{C_k \in \mathbb{R}^{r \times n}, R_k \in \mathbb{R}^{r \times n \times N}} f(C_k, R_k; M_k), \quad (4)$$

where:

$$f(C_k, R_k; M_k) := \frac{1}{2} \sum_{(i,j)} \ell(C_{k,i} R_{k,j} - M_{k,i,j}) + \frac{1}{2} \sum_{(i,j)} \ell(C_{k,i} R_{k,j} - M_{k,i,j}) + \frac{\nu}{2} ||C_k||_F^2 + \frac{\nu}{2} ||R_k||_F^2, \quad (5)$$

where $\ell : \mathbb{R} \to \mathbb{R}$ is the square of the maximum of the two-element set composed of the argument and 0, as described in Section “A Derivation of the Step Size” of the appendix, and $\nu > 0$ is a weight.

Our only further assumption is that we have the element-wise constraints on all elements of the matrixial variable:

**Assumption 1.** For each $(i,j)$ of $M_k$ there is a finite element-wise upper bound $M_{k,i,j}$ and a finite element-wise lower bound $M_{k,i,j}$.

This assumption is satisfied even for any missing values at $ij$ when the measurements lie naturally in a bounded set, e.g., $[0, 255]$ in many computer-vision applications.

**Proposed Algorithms**

In this section, we first present the overall schema of our approach in Algorithm 1. Second, we present Algorithm 2 for on-line inequality-constrained matrix completion, a crucial sub-problem.

**The Overall Schema**

Overall, we interleave the updates to the low-rank model via the inequality-constrained matrix completion, detection of sparse noise, and updating of the inputs to the inequality-constrained matrix completion, which disregards the sparse noise.

At each time step, we acquire new measurements $x_d$ and compute their projection coefficients onto the low-rank subspace as

$$v = \arg \min_{v \in \mathbb{R}^{r \times 1}} ||x_d - vR_{k-1}||_p, \quad (6)$$

where $p$ can be the $1, 2, \infty$ norm, or the 0 pseudo-norm. Since for a very large number of sensors, even solving (6) can be challenging, we subsample $x_d$ by picking only a few sensors uniformly at random. Let $i \in \hat{N}$ be the sampled sensors, with

Input: Initial matrices $(C_0, R_0)$, rank $r$
Output: $(C_k, R_k)$ and events for each $k$

1: for each time $t_k$, $k = 1, 2, \ldots, t_{k+1} - t_k = h$ do
2: acquire new measurements $x_d$
3: subsample $x_d$ uniformly at random to obtain $\tilde{x}_d$
4: compute $\tilde{v}$ via the subsampled projection (7)
5: for each sensor $i$ in parallel do
6: compute residuals $r_i = ||(x_d)_i - (\tilde{v}R_{k-1})_i||$
7: end for
8: compute $\lambda$ as a function of $\{r_i\}$ as described in the appendix
9: compute $T$ as a value at risk of $\{r_i\}$
10: initialise $y$ as a boolean all-False vector of same dimension as $x_d$
11: for each sensor $i$ in parallel do
12: if $r_i < T$ then
13: set $y_i$ to True, as value at sensor $i$ is likely to come from our model
14: add $(x_d)_i$ to $M_k$
15: end if
16: end for
17: compute $(C_k, R_k)$ via Algorithm 2 with rank $r$
18: end for
19: return $(C_k, R_k, y)$

Algorithm 1: Pursuit of low-rank models of time-varying matrices robust to both sparse and measurement noise.

$|\hat{N}| = \hat{N}$. We form a low-dimensional measurement vector $\tilde{x}_d \in \mathbb{R}^{r \times n \hat{N}}$ and solve the subsampled:

$$\tilde{v} = \arg \min_{v \in \mathbb{R}^{r \times 1}} ||\tilde{x}_d - v(R_{k-1})_i||_p, \quad (7)$$

where $(R_{k-1})_i \in \mathbb{R}^{r \times n \hat{N}}$ is the matrix whose columns corresponds to the sensors, which are sampled uniformly at random. Solving (7) yields solutions $\tilde{v}$ such that the norm $||\tilde{v} - \tilde{v}||_p$ is very small, while being considerably less demanding computationally.

Once the projection coefficients $v$ have been computed, we can compute the discrepancy between the measurement $(x_d)_i$ coming from sensor $i$ and our projection (7), $||x_d)_i - (\tilde{v}R_{k-1})_i||_p$, also known as the residual for sensor $i$. We use the residuals in a two-step thresholding procedure inspired by (Malistov 2014). In the first step, we use residuals to compute a coefficient $\lambda > 0$. In the second step, we consider the individual residuals as samples of an empirical distribution, and take the value at risk (VaR) at $\lambda$ as a threshold. We provide details in (Akhirov and Marecek 2019, Akhirov, Marecek, and Simonetto 2018). The test as to whether residual at each sensor is below the threshold results in a binary map, suggesting whether the observation of each sensor is likely to have come from our model or not. For a positive value at $i$ in the map, the measurement $(x_d)_i$ is kept in $M_k$. Otherwise, it is discarded.

**On-line Matrix Completion**

Given $M_k$, we utilise inequality-constrained matrix completion, to estimate the low-rank approximation $(C_k, R_k)$.
of the input matrix considering interval uncertainty sets. Clearly, solving the non-convex problem \( \hat{\delta} \) for non-trivial dimensions of matrix \( M_k \) to a non-trivial accuracy at high-frequency requires careful algorithm design. We propose an algorithm that tracks the low-rank \( R_k \) over time, increasing the accuracy of the solution of \( \hat{\delta} \), while new observations are brought in, and old ones are discarded. In particular, we propose the on-line alternating parallel randomised block-coordinate descent method summarized in Algorithm 2.

For each input \( k \), the previously-found approximate solutions \( (C_{k-1}, R_{k-1}) \), are updated based on the new observation matrix \( M_k \), the correspondingly-derived element-wise lower and upper bounds \( \hat{M}_{k,ij}, \bar{M}_{k,ij} \), and the desired rank \( r \). The update is computed using the alternating least squares (ALS) method, which is based on the observation that while \( f(\hat{\delta}) \) is not convex jointly in \( (C_k, R_k) \), it is convex in \( C_k \) for fixed \( R_k \) and in \( R_k \) for fixed \( C_k \). The update takes the form of a sequence \( \{ (C_k^{T,\tau}, C_k^{T,\tau+1}) \} \) of solutions, which are progressively more accurate. If we could run a large number of iterations of the ALS, we would be in an off-line mode. In the on-line mode, we keep the number of iterations small, and apply the final update based on \( M_k \) at time \( t_k+1 \), when the next observation arrives.

The optimisation in each of the two alternating least-squares problems is based on parallel block-coordinate descent, as reinterpreted by (Nesterov 2012). Notice that in Nesterov’s optimal variant, one requires the the modulus of Lipschitz continuity restricted to the sampled coordinates (Nesterov 2012 Equation 2.4) to compute the step \( \delta \). Considering that the modulus is not known \textit{a priori}, we maintain an estimate \( \hat{W}_{T,\tau} \) of the modulus of Lipschitz continuity restricted to the \( C_{k,\tau} \) sampled, and estimate \( V_{T,\tau} \) of the modulus of Lipschitz continuity restricted to the \( R_{T,\tau} \) sampled. We refer to the appendix for the details of the estimate and to (Nesterov 2012) for a high-level overview.

Overall, when looking at Algorithm 2, notice that there are several nested loops. The counter for the update of the input is \( k \). For each input, we consider factors \( C \) and \( R \) as the optimisation variable alternatingly, with counter \( T \). For each factor, we take a number of block-coordinate descent steps, with the blocks sampled randomly; the counter for the block-coordinate steps is \( \tau \). In particular, in Steps 3–8 of the algorithm, we fix \( R_k^{T,\tau} \), choose a random \( \tilde{r} \) and a random set \( \hat{S}_{\text{row}} \) of rows of \( C_k \), and, in parallel for \( i \in \hat{S}_{\text{row}}, \) update \( C_{k,\tau+1} \) to \( C_{k,\tau} + \delta_{i,k} \), where the step is:

\[
\delta_{i,k} := -\langle \nabla C_k f(C_k^{T,\tau}, R_k^{T,\tau}; M_k), P_{\tilde{r},i} \rangle / W_{T,\tau} \hat{r} \hat{r}, \tag{8}
\]

and \( P_{\tilde{r},i} \) is the \( n \times r \) matrix with 1 in entry \( (i, \tilde{r}) \) and zeros elsewhere. The computation of \( \langle \nabla C_k f(C_k^{T,\tau}, R_k^{T,\tau}; M_k), P_{\tilde{r},j} \rangle \) can be simplified considerably, as explained in in Section “A Derivation of the Step Size” of the appendix.

Likewise, in Steps 9–14, we fix \( C_k^{T,\tau+1} \), choose a \( \tilde{r} \) and a random set \( \hat{S}_{\text{column}} \) of columns of \( R_k \), and, in parallel for \( j \in \hat{S}_{\text{column}}, \) update \( R_{k,\tau+1} \) to \( R_{k,\tau} + \delta_{\tilde{r},j} \), where the step is:

\[
\delta_{\tilde{r},j} := -\langle \nabla R_k f(C_k^{T,\tau+1}, R_k; M_k), P_{\tilde{r},j} \rangle / V_{T,\tau} \hat{r} \hat{r}, \tag{9}
\]

and \( P_{\tilde{r},j} \) is the \( r \times m \) matrix with 1 in entry \( (\tilde{r}, j) \) and zeros elsewhere. Again, the computation of \( \langle \nabla R_k f(C_k^{T,\tau+1}, R_k; M_k), P_{\tilde{r},j} \rangle \) can be simplified.

**Convergence Analysis**

For the off-line inequality-constrained matrix completion problem \( \hat{\delta} \), Marecek, Richtarik, and Takac (2017) proposed an algorithm similar to Algorithm 2 and presented a convergence result, which states that the method is monotonic and, with probability 1, converges to the so-called bistable point, i.e., \( \liminf_{T \to \infty} \| \nabla f(C^*, R^*; M) \| = 0 \), and \( \liminf_{T \to \infty} \| V f(C^*, R^*; M) \| = 0 \). Here, we need to show the rate of convergence to the bistable point and a distance of the bi-stable point to an optimum \( f^* \):

**Theorem 2.** There exists \( \tau > 0 \), such that Algorithm 2 with the initialization to all-zero vector after at most \( T = O(\log \frac{1}{\epsilon}) \) steps has \( f(C^T, R^T) \leq f^* + \epsilon \) with probability 1.

The proof is available online on-line the appendix and should not be surprising, in light of (Bhojanapalli, Neyshabur, and Srebro 2016; Boumal, Voroninski, and Bandeira 2016; Jain and Kar 2017; Boumal, Absil, and Cartis 2018; Bhojanapalli et al. 2018).

Building upon this, we can prove a bound on the error in the on-line regime. In particular, we will show that Algorithm 2 generates a sequence of matrices \( \{(C_k, R_k)\} \) that in the large limit of \( k \to \infty \) guarantees a bounded tracking
error, i.e., $f(C_k, R_k; M_k) \leq f(C^*_k, R^*_k; M_k) + E$. The size of the tracking error $E$ depends on how fast the time-varying matrices change:

**Assumption 3.** The variation of the observation matrix $M_k$ at two subsequent instant $k$ and $k-1$ is so to guarantee that

$$|f(C_k, R_k; M_k) - f(C_k, R_k; M_{k-1})| \leq \epsilon,$$

for all instants $k > 0$.

Now, let us bound the error in tracking, i.e., when $M_k$ changes over time and we run only a limited number of iterations $\tau$ of our algorithm per time step, before obtaining new inputs.

**Theorem 4.** Let Assumptions 1 and 3 hold. Then with probability 1, Algorithm 2 starting from an all-zero matrices generates a sequence of matrices $\{(C_k, R_k)\}$ for which

$$f(C_k, R_k; M_k) - f(C^*_k, R^*_k; M_k) \leq \eta_0 (f(C_{k-1}, R_{k-1}; M_{k-1}) - f(C^*_{k-1}, R^*_{k-1}; M_{k-1})) + \eta_0 \epsilon,$$

where $\eta_0 < 1$ is a constant. In the limit,

$$\limsup_{k \to \infty} \frac{f(C_k, R_k; M_k) - f(C^*_k, R^*_k; M_k)}{C_k} \leq \frac{\eta_0 \epsilon}{1 - \eta_0} =: E.$$

In other words, as time passes, our on-line algorithm generates a sequence of approximately optimal costs that eventually reaches the optimal cost trajectory, up to an asymptotic bound. We bound from above the maximum discrepancy between the approximate optimum and the true one at instant $k$, as $k$ goes to infinity. The convergence to the bound is linear and the rate is $\eta_0$, and depends on the properties of the cost function, while the asymptotic bound depends on how fast the problem is changing over time.

This is a tracking result: we are pursuing a time-varying optimum by a finite number of iterations $\tau$ per time-step. If we could run a large number of iterations per each time step, then we would be back to a off-line case and we would not have a tracking error. This may not, however, be possible in settings, where inputs change faster than one can compute an iteration of the algorithm.

**Experimental Evaluation**

We have implemented Algorithms 1 and 2 in C++, and released the implementation 1 under Apache License 2.0. Based on limited experimentation, we have decided on the use of a time window of $T = 35$, rank $r = 4$, and half-width of the uniform noise $\Delta = 5$. We have used dual simplex from IBM ILOG CPLEX 12.8 as a linear-programming solver for solving solving (7) in Algorithm 1. To initialise the $C_0$ and $R_0$ in Algorithm 1 we have used the matrix completion of Algorithms 2 with 1 epoch per frame for 3 passes on each video (4,000 to 32,000 frames), starting from all-zero matrices. We note that in real-world deployments, such an initialisation may be unnecessary, as the the number of frames processed will render the initial error irrelevant.

First, let us highlight two aspects of the performance of the algorithm. In particular, on the top in Figure 1 we illustrate the effects of the subsampling on the projection (7).

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1 https://github.com/jmarecek/OnlineLowRank

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![Figure 1: Top: Effects of subsampling in the projection (7). Bottom: Performance of Algorithm 2 as a function of the number of epochs per update.](image-url)
at 10 frames per second without parallelisation, which can further improve performance as suggested in Algorithm 2.

We have also conducted a number of experiments on instances from changedetection.net (Goyette et al. 2012), a benchmark often used to test low-rank approaches. There, short videos (1,000 to 9,000 frames) are supplemented with ground-truth information of what is foreground and what is background. These experiments have been run on a single 4-core workstation (Intel Core i7-4800MQ CPU, 16 GB of RAM, RedHat 7.6/64) and results have been deposited on FigShare. In Tables 2 and 3 we summarise the results. In particular, we present the false positive rate (FPR), false negative rate (FNR), specificity, precision, recall, and the geometric mean of the latter two (F1) of our method and 6 other low-rank approaches, which have been used as reference methods recently (Bouwmans, Aybat, and Zahzah 2016). These reference methods are implemented in LRSLibrary (Sobral, Bouwmans, and Zahzah 2015) Bouwmans et al. 2015 and by the original authors of OMoGMF (Meng and Torre 2013) Yong et al. 2018, and have been used with their default settings. Out of these, OMoGMF (Yong et al. 2018) is the most recent and considered to be the most robust. Still, we can improve upon the results of OMoGMF by a considerable margin: the F1 score across the 6 categories is improved by 28% from 0.44643 to 0.57099, for example.

Further details and results are available in the appendix. At http://changedetection.net/ a comparison against four dozen other methods is readily available, although one should like to discount methods tagged as “supervised”, which are trained and tested on one and the same dataset. A further comparison against dozens of other methods is available in (Vaswani et al. 2018).

Conclusions

We have presented a tracking result for time-varying low-rank models of time-varying matrices, robust to both uniformly-distributed measurement noise and arbitrarily-distributed “sparse” noise. This improves upon prior work, as summarised by the recent special issues (Vaswani et al. 2018 Vaswani, Chi, and Bouwmans 2018).

Our analytical guarantees improve upon the state of the art in two ways. First, we provide a bound on the tracking error in estimation of the time-varying low-rank sub-space, rather than a result restricted to the off-line case. Second, we do not make restrictive assumptions on RIP properties, incoherence, identical covariance matrices, independence of all outlier supports, or initialisation. Broadly speaking, such analyses of time-varying non-convex optimisation (Liu et al. 2018) Fattahi et al. 2019 Massicot and Marecek 2019), seems to be an important direction for further research.

In practice, our use of randomised coordinate descent in alternating least-squares seems much better suited to high-volume (high-dimensional, high-frequency) data streams than spectral methods and other alternatives we are aware of. When the matrix $M_k$ does not change quickly, performing a fixed number of iterations within an inexact step upon arrival of a new sample makes it possible to spread the computational load over time, while still recovering a good background model. Also, our algorithm is easy to implement and optimize. It has very few hyper-parameters, and this simplifies tuning. Our results are hence practically relevant.

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Table 2: Results of our Algorithm 2 compared to 6 other approaches on the “baseline” category of http://changedetection.net, evaluated on the 6 performance metrics of (Goyette et al. 2012). For each performance metric, the best result across the presented methods is highlighted in bold.

| Approach / Performance metric                     | Recall   | Specificity | FPR    | FNR    | Precision | F1      |
|---------------------------------------------------|----------|-------------|--------|--------|-----------|---------|
| LRR FastLADMAP (Lin, Liu, and Su 2011)            | 0.74694  | 0.93890     | 0.06021| 0.25306| 0.28039   | 0.36194 |
| MC_GROUSE (Balzano and Wright 2013)              | 0.65640  | 0.89692     | 0.10308| 0.34360| 0.25425   | 0.31495 |
| OMoGMF (Meng and Torre 2013; Yong et al. 2018)   | 0.89943  | 0.98289     | 0.01711| 0.10057| 0.62033   | 0.72611 |
| RPCA_FPCP (Rodriguez and Wohlberg 2013)          | 0.73848  | 0.94733     | 0.05267| 0.26152| 0.29994   | 0.37900 |
| ST_GRASTA (He, Balzano, and Lui 2011)            | 0.45340  | 0.98205     | 0.01795| 0.54660| 0.44009   | 0.42367 |
| TTD_3WD (Oreifej, Li, and Shah 2013)             | 0.61103  | 0.97117     | 0.02883| 0.38897| 0.35557   | 0.40297 |
| Algorithm 2 (w/ Geman-McLure)                    | 0.85684  | 0.99078     | 0.00922| 0.14316| 0.77210   | 0.80254 |
| Algorithm 2 (w/ L1 norm)                         | 0.84561  | 0.99063     | 0.00937| 0.15439| 0.76709   | 0.79421 |

Table 3: Results of our Algorithm 2 compared to 3 other approaches on 6 categories of http://changedetection.net, evaluated on the 6 performance metrics of (Goyette et al. 2012). For each pair of performance metric and category, the best result across the presented methods is highlighted in bold.

| Approach and category / Performance metric         | Recall   | Specificity | FPR    | FNR    | Precision | F1      |
|---------------------------------------------------|----------|-------------|--------|--------|-----------|---------|
| Algorithm 2 (w/ L1 norm)                          |          |             |        |        |           |         |
| badWeather                                        | 0.86589  | 0.98814     | 0.01186| 0.13411| 0.54689   | 0.64618 |
| baseline                                          | 0.84561  | 0.99063     | 0.00937| 0.15439| 0.76709   | 0.79421 |
| cameraJitter                                      | 0.59694  | 0.95928     | 0.04072| 0.40306| 0.55402   | 0.51324 |
| dynamicBackground                                 | 0.46324  | 0.99677     | 0.00323| 0.53676| 0.65511   | 0.49254 |
| nightVideo                                        | 0.83646  | 0.87469     | 0.12531| 0.16354| 0.20992   | 0.29481 |
| shadow                                           | 0.76158  | 0.97612     | 0.02388| 0.23842| 0.64121   | 0.68493 |
| Overall                                           | 0.72829  | 0.96427     | 0.03573| 0.27171| 0.56237   | 0.57099 |
| OMoGMF (Yong et al. 2018)                         |          |             |        |        |           |         |
| badWeather                                        | 0.86871  | 0.98939     | 0.01061| 0.13129| 0.57917   | 0.67214 |
| baseline                                          | 0.89943  | 0.98289     | 0.01711| 0.10057| 0.62033   | 0.72611 |
| cameraJitter                                      | 0.85954  | 0.90739     | 0.09261| 0.14046| 0.30567   | 0.44235 |
| dynamicBackground                                 | 0.87655  | 0.86835     | 0.13617| 0.12345| 0.08601   | 0.15012 |
| nightVideo                                        | 0.75607  | 0.92372     | 0.07628| 0.24393| 0.23522   | 0.31336 |
| shadow                                           | 0.55772  | 0.80276     | 0.03057| 0.27562| 0.40539   | 0.37450 |
| Overall                                           | 0.80300  | 0.91166     | 0.06056| 0.16922| 0.37151   | 0.44643 |
| ST_GRASTA (He, Balzano, and Lui 2011)             |          |             |        |        |           |         |
| badWeather                                        | 0.26555  | 0.98971     | 0.01029| 0.73445| 0.45526   | 0.30498 |
| baseline                                          | 0.45340  | 0.98205     | 0.01795| 0.54660| 0.44009   | 0.42367 |
| cameraJitter                                      | 0.51138  | 0.91313     | 0.08687| 0.48862| 0.23995   | 0.31572 |
| dynamicBackground                                 | 0.41411  | 0.94755     | 0.05245| 0.58589| 0.08732   | 0.13736 |
| nightVideo                                        | 0.42488  | 0.97224     | 0.02776| 0.57512| 0.24957   | 0.28154 |
| shadow                                           | 0.44317  | 0.96861     | 0.03319| 0.55683| 0.42604   | 0.41515 |
| Overall                                           | 0.41875  | 0.96192     | 0.03808| 0.58125| 0.31637   | 0.31307 |
| RPCA_FPCP (Rodriguez and Wohlberg 2013)           |          |             |        |        |           |         |
| badWeather                                        | 0.82546  | 0.84424     | 0.15576| 0.17454| 0.09950   | 0.16687 |
| baseline                                          | 0.73848  | 0.94733     | 0.05267| 0.26152| 0.29994   | 0.37900 |
| cameraJitter                                      | 0.74452  | 0.84143     | 0.15857| 0.25548| 0.18436   | 0.29024 |
| dynamicBackground                                 | 0.69491  | 0.80688     | 0.19312| 0.30509| 0.03928   | 0.07134 |
| nightVideos                                       | 0.79284  | 0.85751     | 0.14249| 0.20716| 0.11797   | 0.19497 |
| shadow                                           | 0.72132  | 0.90454     | 0.09546| 0.27868| 0.26474   | 0.36814 |
| Overall                                           | 0.75292  | 0.86699     | 0.13301| 0.24708| 0.16763   | 0.24509 |
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Proofs

Properties of the Problem

First, let us see that while \( f \) is not convex in both \( C \) and \( R \), it is convex in either \( C \) or \( R \). Jain (Jain and Kar 2017) calls this property marginal convexity: A function \( f(C, R) \) is marginally convex in \( C \), if for every value of \( R \in \mathbb{R}^{r \times n} \), the function \( f(\cdot, R) : \mathbb{R}^{m \times r} \rightarrow \mathbb{R} \) is convex.

Lemma 5 (Marginal Convexity). As continuously differentiable function, \( f : \mathbb{R}^{m \times r} \times \mathbb{R}^{r \times n} \rightarrow \mathbb{R} \) is marginally convex i.e., for every \( C' \), \( C'' \in \mathbb{R}^{m \times r} \), we have

\[
    f(C'', R) \geq f(C', R) + \langle \nabla_x f(C', R) \rangle \cdot (C'' - C'),
\]

where \( \nabla_x f(C', R) \) is the partial gradient of \( f \) with respect to its first variable at the point \((C', R)\), and likewise for \( R \).

Proof. By simple calculus. \( \square \)

Next, let us extend the reasoning of Marecek et al. (Marecek, Richtarik, and Takac 2017) to further properties of the function restricted to only \( C \) or only \( R \). Jain (Jain and Kar 2017) Section 4.4) calls a continuously differentiable function \( f : \mathbb{R}^{m \times r} \times \mathbb{R}^{r \times n} \rightarrow \mathbb{R} \) (uniformly) \( \alpha \)-marginally strongly convex (MSC) in \( C \) if for all \( R \), the function \( f(C, R) \) is \( \alpha \) strongly convex for the constant \( R \). Likewise for (uniformly) \( \beta \)-marginally strongly smooth (MSS) functions. The textbook example (Jain and Kar 2017) Figure 4.1) is \( f : \mathbb{R} \times \mathbb{R} \rightarrow \mathbb{R}, f(x, y) = x \cdot y \). Notice the similarity to the Restricted Isometry Property (RIP) of (Candes 2008).

Lemma 6 (MSC/MSS). There are finite \( \alpha, \beta \), such that the function \( f(\cdot, R) : \mathbb{R}^{m \times r} \rightarrow \mathbb{R} \) is \( \alpha \)-strongly convex and \( \beta \)-strongly smooth, i.e., for every value of \( R \in \mathbb{R}^{r \times n} \), for every \( C', C'' \in \mathbb{R}^{m \times r} \), we have

\[
    \frac{\alpha}{2} \|C'' - C'\|_2^2 \leq f(C'', R) - f(C', R) - \langle g, C'' - C' \rangle \leq \frac{\beta}{2} \|C'' - C'\|_2^2,
\]

where \( g = \nabla_x f(C', R) \) is the partial gradient of \( f \) with respect to its first variable at the point \((C', R)\). Likewise, the function \( f(C, \cdot) : \mathbb{R}^{n \times r} \rightarrow \mathbb{R} \) is \( \alpha' \)-strongly convex and \( \beta' \)-strongly smooth.

Proof of Lemma 6. Notice that \( W_{k, \hat{i}, \hat{r}} \), the modulus of Lipschitz continuity of the gradient of \( f \) restricted to the \( C_{i,r} \) sampled is:

\[
    W_{k, \hat{i}, \hat{r}} = \mu + \sum_{(i,v)} R_{k, i,v}^2
\]

where the superscript denotes squaring, rather than an iteration index, which we omit for brevity. Considering the level set is bounded, \( W_{k, \hat{i}, \hat{r}} \) is bounded and we have the result. Similarly \( V_{\hat{i}, \hat{r}} \), the modulus of Lipschitz continuity of the gradient of \( f \) restricted to the \( R_{k, i,r} \) is:

\[
    V_{k, \hat{i}, \hat{r}} = \mu + \sum_{(i,v)} C_{k, i,v}^2
\]

where again, the superscript denotes squaring, rather than an iteration index. \( \square \)

Next, let us consider some more definitions of (Jain and Kar 2017). For any \( R \), we say that \( C \) is a marginally optimal coordinate with respect to \( R \), and use the shorthand \( C \in \text{OPT}_f(R) \), if \( f(C, R) \leq f(C, \hat{R}) \) for all \( C \). Similarly for any \( C, \hat{R} \in \text{OPT}_f(C) \) if \( \hat{R} \) is a marginally optimal coordinate with respect to \( C \). Then:

Definition 7 (Bistable Point of (Jain and Kar 2017)). Given a function \( f \) over two variables constrained within the sets \( X, Y \) respectively, a point \((C, R) \in X \times Y \) is considered a bistable point if \( y \in \text{OPT}_f(C) \) and \( x \in \text{OPT}_f(y) \) i.e., both coordinates are marginally optimal with respect to each other.

Lemma 8 (Jain et al. (Jain and Kar 2017)). A point \((C, R) \) is bistable with respect to a continuously differentiable function \( f : \mathbb{R}^{m \times r} \times \mathbb{R}^{r \times n} \) that is marginally convex in both its variables if and only if \( \nabla f(C, R) = 0 \).

Proof of Lemma 8. Notice that each element of the matrix is bounded both from above and from below. The level sets are hence bounded, whereby we obtain the result. \( \square \)

Then, we can restate Theorem 1 of (Marecek, Richtarik, and Takac 2017):

Theorem 9 (Based on Theorem 1 in Marecek et al. (Marecek, Richtarik, and Takac 2017)). For any \( \tau > 0 \) and \( \hat{S}_{\text{row}}, \hat{S}_{\text{column}} \) sampled uniformly at random, the limit point \( \liminf_{T \to \infty} (C_k^{\tau}, R_k^{\tau}) \) of Algorithm 2 is bistable with probability 1.

The proof follows that of Theorem 1 in (Marecek, Richtarik, and Takac 2017). There, however, the analysis of (Marecek, Richtarik, and Takac 2017) ends.
The Limit Point
Next, consider further properties of the limit point under the assumptions above. To do so, we present some more definitions of Jain (Jain and Kar 2017).

Definition 10 (Robust Bistability Property of (Jain and Kar 2017)). A function $f : \mathbb{R}^{m \times r} \times \mathbb{R}^{r \times n} \to \mathbb{R}$ satisfies the $C$-robust bistability property if for some $C > 0$, for every $(C, R) \in \mathbb{R}^{m \times r} \times \mathbb{R}^{r \times n}$, $\hat{R} \in \text{OPT}_f(C)$ and $\hat{C} \in \text{OPT}_f(R)$, we have

$$f(C, R*) + f(C*, \hat{R}) - 2f* \leq C \cdot \left(2f(C, \hat{R}) - f(C, \hat{R}) - f(\hat{C}, \hat{R})\right).$$

Subsequently:

Lemma 11. Under Assumption [1] there exists a finite $C > 0$, such that the function $f(4)$ satisfies the $C$-robust bistability property.

Much more detailed results, bounding the constant $C$, are available in many regimes, e.g., when each element of the matrix is sampled with a probability larger than a certain instance-specific $p$ from a certain ensemble (Cheng and Ge 2018), and more generally when one allows from a certain smoothing (Bhojanapalli, Neyshabur, and Srebro 2016; Bhojanapalli et al. 2018). Further, one can use the results of (Park et al. 2017) to prove its satisfaction under the Restricted Isometry Property (RIP) of Candès (2008).

Next, let us state a technical lemma:

Lemma 12 (Based on Lemma 4.4 of (Jain and Kar 2017)). Under Assumption [1], for any $(C, R) \in \mathbb{R}^{m \times r} \times \mathbb{R}^{r \times n}$, $\hat{R} \in \text{OPT}_f(C)$ and $\hat{C} \in \text{OPT}_f(R)$,

$$\|C - C*\|_2^2 + \|R - R*\|_2^2 \leq \frac{C\beta}{\alpha} \left(\|C - \hat{C}\|_2^2 + \|R - \hat{R}\|_2^2\right)$$

Proof of Lemma 12 Notice that $f$ is $\alpha$-MSC, $\beta$-MSS in both $C$ and $R$, as shown in Lemma 5 and 6. From Lemma 6,

$$f(C, R*) + f(C*, \hat{R}) \geq 2f* + \frac{\alpha}{2} \left(\|C - C*\|_2^2 + \|\hat{R} - C*\|_2^2\right)$$

$$2f(C, R) \leq f(C, \hat{R}) + f(\hat{C}, R) + \frac{\beta}{2} \left(\|C - \hat{C}\|_2^2 + \|R - \hat{R}\|_2^2\right)$$

Applying robust bistability of Lemma 11 then proves the result. □

Using Lemma 12 we can present a bound on the limit point and the rate of convergence to it, i.e., prove Theorem 2 which we restate here for convenience:

Theorem 13. There exists $\tau > 0$, such that Algorithm 2 with the initialization to all-zero vector after at most $T = O(\frac{1}{\epsilon})$ steps has $f(C^T, R^T) \leq f* + \epsilon$ with probability $\tau$.

Proof. We follow (Jain and Kar 2017) and use $\phi(k) = f(C(k), R(k)) - f*$ as the potential function. The $\tau$ we require depends on the cardinality of $\hat{S}_{\text{row}}$, $\hat{S}_{\text{column}}$, and the model of computation, but should be large enough for marginal optimisation, i.e., $\nabla_C f(C*, R*) = 0$. Then, Lemma 6 assures:

$$f(C^{(k+1)}, R*) - f(C*, R*) \leq \frac{\beta}{2} \|C^{(k+1)} - C*\|_2^2$$

Further, considering $R^{(k+1)} \in \text{OPT}_f(C^{(k+1)})$, we have:

$$\phi^{(k+1)} = f(C^{(k+1)}, R^{(k+1)}) - f* \leq f(C^{(k+1)}, R*) - f* \leq \frac{\beta}{2} \|C^{(k+1)} - C*\|_2^2$$

Again, considering $\nabla_C f(C^{(k+1)}, R^{(k)}) = 0$ for large-enough $\hat{S}_{\text{row}},$

$$f(C^{(k)}, R^{(k)}) \geq f(C^{(k+1)}, R^{(k)}) + \frac{\alpha}{2} \|C^{(k+1)} - C^{(k)}\|_2^2 \geq f(C^{(k+1)}, R^{(k+1)}) + \frac{\alpha}{2} \|C^{(k+1)} - C^{(k)}\|_2^2$$

and consequently

$$\phi^{(k)} - \phi^{(k+1)} \geq \frac{\alpha}{2} \|C^{(k+1)} - C^{(k)}\|_2^2.$$
Applying Lemma 12,
\[
\|C^{(k)} - C^*\|^2_2 \leq \|C^{(k)} - C^*\|^2_2 + \|R^{(k)} - R^*\|^2_2 \\
\leq \frac{C\beta}{\alpha} \|C^{(k)} - C^{(k+1)}\|^2_2.
\]
Using \((a + b)^2 \leq 2(a^2 + b^2),\)
\[
\Phi^{(k+1)} \leq \frac{\beta}{2} \|C^{(k+1)} - C^*\|^2_2 \\
\leq \beta \left( \|C^{(k+1)} - C^{(k)}\|^2_2 + \|C^{(k)} - C^*\|^2_2 \right) \\
\leq \beta(1 + C\kappa) \|C^{(k+1)} - C^{(k)}\|^2_2 \\
\leq 2\kappa(1 + C\kappa) \left( \Phi^{(k)} - \Phi^{(k+1)} \right).
\]
Finally, by simple algebra,
\[
\Phi^{(k+1)} \leq \eta_0 \cdot \Phi^{(k)},
\]
where
\[
\eta_0 = \frac{2\kappa(1 + C\kappa)}{1 + 2\kappa(1 + C\kappa)} < 1.
\]

Finally:

Proof of Theorem 4. The proof follows from Theorem 2 by invoking the triangle inequality and the sum of a geometric series. In particular, due to Theorem 2, one has for each \(k\)
\[
f(C_k, R_k; M_k) - f(C^*_k, R^*_k; M_k) \leq \eta_0(f(C_{k-1}, R_{k-1}; M_{k-1}) - f(C^*_k, R^*_k; M_k)).
\]
By summing and subtracting \(\eta_0 f(C_{k-1}, R_{k-1}; M_{k-1})\) to the right-hand-side and putting without loss of generality \(f(C^*_k, R^*_k; M_{k-1}) = f(C^*_k, R^*_k; M_{k-1})\),
\[
f(C_k, R_k; M_k) - f(C^*_k, R^*_k; M_k) \leq \eta_0(f(C_{k-1}, R_{k-1}; M_{k-1}) - f(C^*_k, R^*_k; M_{k-1}) + \eta_0(f(C_{k-1}, R_{k-1}; M_{k-1}) - f(C^*_k, R^*_k; M_{k-1}) + e),
\]
and by using Assumption 3
\[
f(C_k, R_k; M_k) - f(C^*_k, R^*_k; M_k) \leq \eta_0(f(C_{k-1}, R_{k-1}; M_{k-1}) - f(C^*_k, R^*_k; M_{k-1})) + \eta_0 e.
\]
By summation of geometric series, the claim is proven.
Details of the Thresholding

As suggested in the main body of the text, we start by looking for the best linear combination \(c\) that minimizes difference in \(L_1\):

\[
e_{opt} = \arg \min_{c} \|cR - f\|_1 = \arg \min_{c} \sum_{i=1}^{N} |(cR)_{i} - f_{i}|,\]  

(21)

where \(c\) is a \(1 \times \text{rank}\) vector, \(f\) is a 2D image flattened into \(1 \times N\) vector, and \((cR)_{i}\) is the scalar result of multiplication between vector \(c\) and \(i\)-th column of matrix \(R\). Due to the robust property of \(L_1\) norm, the formulation (21) provides a close approximation of the new frame at the majority of stationary (background) points, while leaving residuals at the “moving” (foreground) points relatively high. By introducing the additional variables \(m_i\): \(|(cR)_{i} - f_{i}| \leq m_i\), for all \(i = 1, N\), the optimization problem can be reformulated as a linear program:

\[
\begin{align*}
\text{minimize:} & \quad \sum_{i=1}^{N} m_i \\
\text{subject to:} & \quad 0 \leq m_i < +\infty, \\
& \quad -\infty < (cR)_{i} - m_i \leq f_{i}, \\
& \quad f_{i} \leq (cR)_{i} + m_i < +\infty.
\end{align*}
\]  

(22)

Alternatively, one can consider the robust Geman-McLure function \(\rho(r, \sigma) = r^2/(r^2 + \sigma^2)\) as featured in (Sawhney and Ayer 1996), where parameter \(\sigma\) is estimated from the distribution of residuals over the course of optimization

\[
e_{opt} = \arg \min_{c} \sum_{i=1}^{N} \rho ((cR)_{i} - f_{i}).
\]  

(23)

In practice, both (21) and (23) produce results of similar quality, with a slightly better statistical performance of (21) at a minor additional expense in terms of run-time, compared to the use of gradient methods (Sawhney and Ayer 1996) in minimisation of (23).

After the optimal linear combination \(e_{opt}\) has been obtained in (21), the next step is to compute residuals \(r_i = |(cR)_{i} - f_{i}|\) and threshold them into those generated by the low-rank model, \(r_i < T\), and the remainder, \(r_i \geq T\), where \(T\) is some threshold. Thresholding for background subtraction is a vast area by itself. Although locally adapted threshold may work best, it is quite common to choose a single threshold for each frame. We follow the same practice: As often (Malistov 2014, Akhriev and Marecek 2019) in Computer Vision, we seek a threshold of the highest sensitivity, when isolated points "just" show up. In particular, we seek a threshold such that a certain fraction (0.0025) of 3 \(\times\) 3 contiguous patches have 1 or 2 pixels exceeding the threshold, as suggested in Figure 2. To explain this in detail, consider the RGB colour images, where the point-wise 2D residual map is computed as follows:

\[
r_i = \left| R_i^{(f)} - R_i^{(b)} \right| + \left| G_i^{(f)} - G_i^{(b)} \right| + \left| B_i^{(f)} - B_i^{(b)} \right|,
\]

where subscripts \(f\) and \(b\) stands for current frame and background respectively, and index \(i\) enumerates image pixels. Other metrics like Euclidean one are also possible. We accumulate so called histogram of thresholds by analysing each point in the residual map. There are several how residual value at the central point of relates to its neighbour.

Let us consider one example. Suppose, the central value in the largest one \(v_1\) and we pick up the second \(v_2\) and the third \(v_3\) largest ones from the \(3 \times 3\) vicinity, \(v_2 \leq v_2 \leq v_3\), and all the values are integral as usual for images. If a threshold happens in the interval \([v_3 + 1 \ldots v_1]\) then one of the patterns depicted on Figure 2 will show up after thresholding. As such, this particular point “votes” for the range \([v_3 + 1 \ldots v_1]\) in the histogram of thresholds, which means we increment counters in the bins \(v_3 + 1\) to \(v_1\). Repeating the process for all the points, we arrive to the histogram of thresholds as shown in Figure 3. The region around the mode of the histogram (50% of its area), outlined by yellow margins on Figure 3 mostly contains noise. We start search for the optimum threshold from the right margin to the right until the value of histogram bin is less then 0.0025 \(\times N\), where \(N\) is the number of pixels. We found experimentally that the fraction 0.0025 works the best, although its value can be varied without drastic effect.

For another example of use of similar thresholding techniques, please see (Akhriev and Marecek 2019).

![Figure 2: The configurations of the 3 \(\times\) 3 contiguous patches, whose fraction within all the 3 \(\times\) 3 contiguous patches is sought.](image-url)
Figure 3: A histogram of residuals. The histogram was truncated from the original $3\cdot255$ residuals to allow for some clarity of presentation. In green, there is the middle of the least-width interval representing half of the mass. In yellow, there are the end-points of the interval. In red, the “optimal” threshold we use.
A Derivation of the Step Size

Minimisation of the objective function in $C_{k,ir}$

\[
f(C_k, R_k) = \frac{\mu}{2} \sum_{i,j} (C_{k,ij}^2 + R_{k,ij}^2) + \frac{1}{2} \sum_{C_{k,i} R_{k,j} < M_{k,ij}} (C_{k,i} R_{k,j} - M_{k,ij})^2 + \frac{1}{2} \sum_{C_{k,i} R_{k,j} > M_{k,ij}} (C_{k,i} R_{k,j} - M_{k,ij})^2.
\]  

(24)

\[
\frac{\partial f}{\partial C_{k,ir}} = \mu C_{k,ir} + \sum_{j : C_{k,i} R_{k,j} < M_{k,ij}} (C_{k,i} R_{k,j} - M_{k,ij}) R_{k,rj} + \sum_{j : C_{k,i} R_{k,j} > M_{k,ij}} (C_{k,i} R_{k,j} - M_{k,ij}) R_{k,rj}.
\]  

(25)

\[
W_{ir} \triangleq \mu + \sum_{j : C_{k,i} R_{k,j} < M_{k,ij}} R_{k,rj}^2 + \sum_{j : C_{k,i} R_{k,j} > M_{k,ij}} R_{k,rj}^2.
\]  

(26)

\[
\delta = - \frac{\partial f}{\partial C_{k,ir}} / W_{k,ir}.
\]  

(27)

\[
C_{k,ir} \leftarrow C_{k,ir} + \delta.
\]  

(28)

\[
A_{k,ij} \leftarrow A_{k,ij} + \delta R_{k,rj} \quad \forall j.
\]  

(29)
Minimisation of the objective function in $R_{k,rj}$

\[
f(C_k, R_k) = \frac{\mu}{2} \sum_{i,j} \left( C_{k,ij}^2 + R_{k,ij}^2 \right) + \frac{1}{2} \sum_{i : C_{k,i} < M_{k,ij}} (C_{k,i} - M_{k,ij})^2 + \frac{1}{2} \sum_{i : C_{k,i} > M_{k,ij}} (C_{k,i} - M_{k,ij})^2.
\] (30)

\[
\frac{\partial f}{\partial R_{k,rj}} = \mu R_{k,rj} + \sum_{i : C_{k,i} < M_{k,ij}} (C_{k,i} - M_{k,ij}) C_{k,ir} + \sum_{i : C_{k,i} > M_{k,ij}} (C_{k,i} - M_{k,ij}) C_{k,ir}.
\] (31)

\[
V_{k,rj} \triangleq \mu + \sum_{i : C_{k,i} < M_{k,ij}} C_{k,ir}^2 + \sum_{i : C_{k,i} > M_{k,ij}} C_{k,ir}^2.
\] (32)

\[
\delta = -\frac{\partial f}{\partial R_{k,rj}} / V_{k,rj}.
\] (33)

\[
R_{k,rj} \leftarrow R_{k,rj} + \delta.
\] (34)

\[
A_{k,ij} \leftarrow A_{k,ij} + \delta C_{k,ir} \quad \forall i.
\] (35)
In Table 4, we present the overall results on changedetection.net as the average over all the frames of a video, with a standard deviation in parentheses. First, we present MS-SSIM of (Vang, Simoncelli, and Bovik 2003), a well-known measure of similarity of the background of each frame to our rank-4 estimate thereof, which is also known as the multiscale structural similarity for image quality. There, our estimates perform rather well, with the exception of videos featuring dynamic backgrounds such as waves and reflections of sun light on water, where the low-rank model is not updated often enough to capture all of the rapid changes. Next, we present the F1 score, which is the harmonic mean of precision and recall and which we used the code provided by CDnet to evaluate against the ground truth. We should like to stress that the F1 score depends on thresholding method, which is quite simple in our current implementation and could be improved. Finally, a number of modern methods including the top three in the CDnet ranking as of May 2018 are “supervised”, in the sense that they derive megabytes of a model from the test set and then apply the model to the test set, which constitutes “double dipping”. With these caveats in mind, the performance seems rather respectable.

| Video sequence                        | MS-SSIM | F1-score | Recall | Precision |
|---------------------------------------|---------|----------|--------|-----------|
| badWeather/blizzard                   | 0.990   | 0.752    | 0.901  | 0.675     |
| badWeather/skating                    | 0.980   | 0.872    | 0.890  | 0.891     |
| badWeather/snowFall                   | 0.976   | 0.601    | 0.832  | 0.505     |
| badWeather/wetSnow                    | 0.979   | 0.446    | 0.838  | 0.356     |
| baseline/PETS2006                     | 0.983   | 0.769    | 0.963  | 0.655     |
| baseline/highway                      | 0.946   | 0.886    | 0.848  | 0.934     |
| baseline/office                       | 0.959   | 0.652    | 0.647  | 0.745     |
| baseline/pedestrians                  | 0.988   | 0.930    | 0.989  | 0.887     |
| dynamicBackground/boats               | 0.794   | 0.316    | 0.209  | 0.890     |
| dynamicBackground/canoe               | 0.758   | 0.692    | 0.561  | 0.994     |
| dynamicBackground/fall                | 0.824   | 0.274    | 0.291  | 0.430     |
| dynamicBackground/fountain01          | 0.919   | 0.245    | 0.336  | 0.208     |
| dynamicBackground/fountain02          | 0.957   | 0.785    | 0.702  | 0.925     |
| dynamicBackground/overpass            | 0.935   | 0.644    | 0.584  | 0.778     |
| intermittentObjectMotion/abandonedBox | 0.997   | 0.563    | 0.505  | 0.724     |
| intermittentObjectMotion/parking      | 0.945   | 0.230    | 0.190  | 0.868     |
| intermittentObjectMotion/sofa         | 0.979   | 0.518    | 0.510  | 0.585     |
| intermittentObjectMotion/steetLight   | 0.999   | 0.339    | 0.294  | 0.756     |
| intermittentObjectMotion/tramstop     | 0.977   | 0.393    | 0.293  | 0.727     |
| intermittentObjectMotion/winterDriveway| 0.970 | 0.394    | 0.930  | 0.286     |
| lowFrameRate/port_0.17fps             | 0.988   | 0.223    | 0.557  | 0.187     |
| lowFrameRate/tramCrossroad_1fps       | 0.995   | 0.758    | 0.934  | 0.663     |
| lowFrameRate/tunnelExit_0.35fps       | 0.979   | 0.628    | 0.836  | 0.564     |
| lowFrameRate/turnpike_0.5fps          | 0.967   | 0.736    | 0.639  | 0.947     |
| nightVideos/bridgeEntry              | 0.980   | 0.098    | 0.977  | 0.053     |
| nightVideos/busyBoulard              | 0.995   | 0.304    | 0.623  | 0.259     |
| nightVideos/fluuidHighway            | 0.935   | 0.103    | 0.946  | 0.059     |
| nightVideos/streetCornerAtNight      | 0.985   | 0.281    | 0.838  | 0.188     |
| nightVideos/tramStation              | 0.986   | 0.750    | 0.895  | 0.668     |
| nightVideos/winterStreet             | 0.955   | 0.202    | 0.946  | 0.119     |
| shadow/backdoor                      | 0.984   | 0.889    | 0.924  | 0.874     |
| shadow/bungalows                     | 0.949   | 0.553    | 0.620  | 0.670     |
| shadow/busStation                    | 0.960   | 0.733    | 0.765  | 0.775     |
| shadow/copyMachine                   | 0.942   | 0.571    | 0.758  | 0.528     |
| shadow/cubicle                       | 0.983   | 0.696    | 0.757  | 0.705     |
| shadow/peopleInShade                 | 0.968   | 0.823    | 0.992  | 0.754     |

Table 4: Results on changedetection.net.
In Table 5, we present a comparison similar to Table 3 except that results of Algorithm 2 are obtained by using the smooth Geman-McLure loss function instead of subsampling with the non-smooth L1 norm.

Table 5: Further results on [http://changedetection.net](http://changedetection.net).

| Method / Category | Recall  | Specificity | FPR   | FNR   | Precision | F1    |
|-------------------|---------|-------------|-------|-------|-----------|-------|
| **Algorithm 2 (w/ Geman-McLure):** |         |             |       |       |           |       |
| badWeather        | 0.86733 | 0.98695     | 0.01305 | 0.13267 | 0.52229   | 0.62347 |
| baseline          | 0.85684 | **0.99078** | **0.00922** | 0.14316 | **0.77210** | **0.80254** |
| cameraJitter      | 0.59669 | **0.95909** | **0.04091** | 0.40331 | **0.56832** | **0.51461** |
| dynamicBackground | 0.46994 | **0.96272** | **0.00373** | 0.53006 | **0.63826** | **0.49071** |
| nightVideo        | **0.83068** | 0.87444     | 0.12556 | **0.16932** | 0.20758   | 0.29115 |
| shadow            | **0.76715** | **0.97409** | 0.02591 | 0.23285 | **0.61857** | **0.66833** |
| Overall           | 0.73144 | **0.96360** | **0.03640** | 0.26856 | **0.55452** | **0.56514** |
| **OMoGMF (Yong et al. 2018):** |         |             |       |       |           |       |
| badWeather        | **0.86871** | 0.98939     | 0.01061 | 0.13129 | **0.57917** | **0.67214** |
| baseline          | **0.89943** | 0.98289     | 0.01711 | **0.10057** | 0.62033   | 0.72611 |
| cameraJitter      | **0.85954** | 0.90739     | 0.09261 | **0.14046** | 0.30567   | 0.44235 |
| dynamicBackground | **0.87655** | 0.86383     | 0.13617 | **0.12345** | 0.08601   | 0.15012 |
| nightVideo        | 0.75607 | 0.92372     | 0.07628 | 0.24393 | 0.23252   | 0.31336 |
| shadow            | 0.55772 | 0.80276     | 0.03057 | 0.27562 | 0.40539   | **0.37450** |
| Overall           | **0.80300** | 0.91166     | 0.06056 | **0.16922** | 0.37151   | 0.44643 |
| **STGRASTA (He, Balzano, and Lui 2011):** |         |             |       |       |           |       |
| badWeather        | 0.26555 | **0.98971** | **0.01029** | 0.73445 | 0.45526   | 0.30498 |
| baseline          | 0.45340 | 0.98205     | 0.01795 | 0.54660 | 0.44009   | 0.42367 |
| cameraJitter      | 0.51138 | 0.91313     | 0.08687 | 0.48862 | 0.23995   | 0.31572 |
| dynamicBackground | 0.41411 | 0.94755     | 0.05245 | 0.58589 | 0.08732   | 0.13736 |
| nightVideo        | 0.42488 | **0.97224** | **0.02776** | 0.57512 | **0.24957** | 0.28154 |
| shadow            | 0.44317 | 0.96681     | 0.03319 | 0.55683 | 0.42604   | 0.41515 |
| Overall           | 0.41875 | 0.96192     | 0.03808 | 0.58125 | 0.31637   | 0.31307 |
| **RPCA_FPCP (Rodriguez and Wohlberg 2013):** |         |             |       |       |           |       |
| badWeather        | 0.82546 | 0.84424     | 0.15576 | 0.17454 | 0.09950   | 0.16687 |
| baseline          | 0.73848 | 0.94733     | 0.05267 | 0.26152 | 0.29994   | 0.37900 |
| cameraJitter      | 0.74452 | 0.84143     | 0.15857 | 0.25548 | 0.18436   | 0.29024 |
| dynamicBackground | 0.69491 | 0.80688     | 0.19312 | 0.30509 | 0.03928   | 0.07134 |
| nightVideo        | 0.79284 | 0.85751     | 0.14249 | 0.20716 | 0.11797   | 0.19497 |
| shadow            | 0.72132 | 0.90454     | 0.09546 | 0.27868 | 0.26474   | 0.36814 |
| Overall           | 0.75292 | 0.86699     | 0.13301 | 0.24708 | 0.16763   | 0.24509 |
In Table 6, we present a comparison of the mean run-time per frame of the methods discussed in the paper. Notice that this corresponds to a speed-up of up to the factor of 103.

Table 6: Mean processing time per input frame (in seconds) on the “baseline/highway” video-sequence from http://changedetection.net. Note, our implementation does not use any parallelisation at the moment. This was done on purpose to run on a machine serving multiple cameras simultaneously.

| Method                    | Mean time per frame |
|---------------------------|---------------------|
| LRR, FastLADMAP (Lin, Liu, and Su 2011) | 4.611               |
| MC_GROUSE (Balzano and Wright 2013)       | 10.621              |
| OMoGMF (Meng and Torre 2013; Yong et al. 2018) | 0.123               |
| RPCA_FPCP (Rodriguez and Wohlberg 2013)   | 0.504               |
| ST_GRASTA (He, Balzano, and Liu 2011)      | 3.266               |
| TTD_3WD (Oreifej, Li, and Shah 2013)       | 10.343              |
| Algorithm 2 (w/ Geman-McLure)              | 0.103               |
| Algorithm 2 (w/ $L_1$ norm)                | 0.194               |
Figure 4: One snapshot from the video baseline/highway (from the top left, clock-wise): one frame of the original video, our estimate of the background, our residuals prior to thresholding, the ground truth, an exponential smoothing of all frames prior to the current one with smoothing factor of $1/35$, and finally, our Boolean map obtained by thresholding residuals.