Rates of readmission and death associated with leaving hospital against medical advice: a population-based study

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Abstract

Background: Leaving hospital against medical advice may have adverse consequences. Previous studies have been limited by evaluating specific types of patients, small sample sizes and incomplete determination of outcomes. We hypothesized that leaving hospital against medical advice would be associated with increases in subsequent readmission and death.

Methods: In a population-based analysis involving all adults admitted to hospital and discharged alive in Manitoba from Apr. 1, 1990, to Feb. 28, 2009, we evaluated all-cause 90-day mortality and 30-day hospital readmission. We used multivariable regression, adjusted for age, sex, socioeconomic status, year of hospital admission, patient comorbidities, hospital diagnosis, past frequency of admission to hospital, having previously left hospital against medical advice and data clustering (patients with multiple admissions). For readmission, we assessed both between-person and within-person effects of leaving hospital against medical advice.

Results: Leaving against medical advice occurred in 21 417 of 1 916 104 index hospital admissions (1.1%), and was associated with higher adjusted rates of 90-day mortality (odds ratio [OR] 2.51, 95% confidence interval [CI] 2.18–2.89), and 30-day hospital readmission (within-person OR 2.10, CI 1.99–2.21; between-person OR 3.04, CI 2.79–3.30). In our additional analyses, elevated rates of readmission and death associated with leaving against medical advice were manifest within 1 week and persisted for at least 180 days after discharge.

Interpretation: Adults who left the hospital against medical advice had higher rates of hospital readmission and death. The persistence of these effects suggests that they are not solely a result of incomplete treatment of acute illness. Interventions aimed at reducing these effects may need to include longitudinal interventions extending beyond admission to hospital.

Methods

Study design
This retrospective cohort study used population-based administrative data from Manitoba, a province with a population of 1.2 million people. These data are collected by the publicly funded health insurance system and housed at the Manitoba Centre for Health Policy. The data contain information on the use of health services for all provincial residents and are linked to census-based socioeconomic information and vital statistics. Our study was approved by the University of Manitoba Health Research Ethics Board and the Manitoba Health Information Privacy Committee.

Study population
We included all residents of Manitoba 18 years of age or older who were discharged alive from acute care hospitals after an unscheduled admission from Apr. 1, 1990, to Feb. 28, 2009. Because
patients can undergo interhospital transfer within a single episode of hospital care, we combined such hospital abstracts to construct complete admissions using previously described criteria, with the modification that 2 abstracts were considered to belong to separate admissions if the earlier abstract indicated the patient had left against medical advice.

We designated discharges from hospital as against medical advice (AMA status) and not against medical advice (non-AMA status) according to whether the final hospital abstract from the admission contained the specific discharge separation code for having left against medical advice. We performed external validation of this coding in 291 hospital abstracts where true AMA status was determined from the nursing and physician progress notes in the hospital charts. These abstracts were randomly chosen, in a ratio of about 1:2 (AMA:non-AMA), from a database used for quality-improvement purposes of all admissions to internal medicine wards in 4 hospitals. All 198 discharges that were not patients leaving against medical advice were correctly identified in the abstracts, giving a specificity of 100% (95% confidence interval [CI] 98.2%–100%). However, only 81 of 93 AMA discharges were correctly coded in the abstracts, giving a sensitivity of 87% (95% CI = 78.5%–93.2%).

Outcome measures

Our primary outcomes were unscheduled readmission to hospital within 30 days of discharge and death within 90 days of discharge.

Statistical analysis

To assess their association with leaving against medical advice, we used multivariable regression models to adjust for potential confounding variables.

We categorized year of admission to achieve equality of admissions as 1990–1993, 1994–1998, 1999–2003 and 2004–2009. We identified the 31 comorbid conditions described by Elixhauser and colleagues and Quan and colleagues, including diagnoses listed during the index admission and admissions up to 1 year before the index admission.

We collapsed the categories of diabetes with and without chronic complications, because this distinction was unclear in coding in Manitoba before 2006.

As a surrogate measure of socioeconomic status, we quantified average household income in the postal code of residence, divided into separate quintiles for rural and urban residents. People living where the Canadian census does not track incomes formed an eleventh category referred to as “not calculated”; most of these patients resided in nursing homes or other long-term care facilities, with a smaller portion living in penitentiaries.

We identified diagnosis as the most responsible hospital diagnosis, defined as the condition responsible for most of the hospital stay, and grouped diagnoses into the 18 main chapter headings of the clinical modification of the International Classification of Diseases, 9th revision.

We evaluated 30-day hospital readmission after all unscheduled admissions to hospital from which patients left alive, referring to this unit of analysis as the index admission. We adjusted for age, sex, year of index admission, comorbid conditions, socioeconomic status, diagnosis at index admission, length of stay, the number of admissions to hospital in the 5 years before the index admission and the existence of any AMA discharges in the 5 years before the index admission.

To account for data clustering, where many patients had multiple admissions over the study period, we used general estimating equations with an exchangeable correlation structure and robust standard errors. When dealing with such data, it is important to note that leaving against medical advice may have different within-person and between-person influences on hospital readmission. Within-person effects indicate the difference in readmission risk after an AMA discharge and a non-AMA discharge for the same person; between-person effects indicate the difference in readmission risk after an AMA discharge for one person and a non-AMA discharge for a different patient with otherwise identical characteristics. We included separate variables representing these 2 effects, as described in Appendix 1 (available at www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.130029/-/DC1). We verified the absence of important multicollinearity among independent variables using variance inflation factors, which averaged 1.3 and had a maximum value of 3.3. To assess the robustness of our results, we performed 3 sensitivity analyses of the logistic regression modelling of 30-day readmissions by using an autoregressive instead of exchangeable correlation structure in the general estimating equations model, by truncating each patient’s included admissions after the first occurrence of leaving against medical advice while omitting the previous AMA discharge covariate and by including only a single randomly chosen index admission per patient, which assessed only the between-person effect.

Accurately assessing the association between leaving against medical advice and 90-day mortality necessitates addressing that patients can have multiple admissions to hospital but only die once. For this purpose, we performed adjusted, matched, case–control analysis, with patients as
Table 1: Characteristics and outcomes of index admissions to hospital used in the analysis of hospital readmission

| Variable†                          | Discharge status of index admission, no. (%)| p value‡         |
|-----------------------------------|---------------------------------------------|------------------|
|                                   | AMA n = 21 417                              | Non-AMA n = 1 894 687 |
| Any AMA discharges in previous 5 yr | 5 076 (23.7)                                | 42 537 (2.3)      | < 0.001 |
| No. of admissions to hospital in previous 5 yr, mean ± SD | 3.8 (6.4)                                  | 2.3 (4.0)         | < 0.001 |
| Median (IQR)                      | 2 (1–4)                                     | 1 (0–3)           | < 0.001 |
| Age, yr, mean ± SD                | 42.7 (18.0)                                 | 54.1 (22.2)       | < 0.001 |
| Male sex                          | 10 789 (50.4)                               | 708 491 (37.4)    | < 0.001 |
| Average household income quintile |                                             |                  |
| Urban 1 (lowest urban stratum)    | 5 831 (27.2)                                | 257 035 (13.6)    | < 0.001 |
| Urban 3                           | 1 557 (7.3)                                 | 185 133 (9.8)     |        |
| Urban 5 (highest urban stratum)   | 598 (2.8)                                   | 132 683 (7.0)     |        |
| Rural 1 (lowest rural stratum)    | 3 864 (18.0)                                | 227 971 (12.0)    |        |
| Rural 3                           | 1 534 (7.2)                                 | 194 665 (10.3)    |        |
| Rural 5 (highest rural stratum)   | 1 058 (4.9)                                 | 123 465 (6.5)     |        |
| Not calculated                    | 272 (1.3)                                   | 47 764 (2.5)      |        |
| Comorbid conditions               |                                             |                  |
| Alcohol abuse                     | 6 010 (28.1)                                | 74 756 (4.0)      | < 0.001 |
| Diabetes                          | 3 341 (15.6)                                | 227 715 (12.0)    | < 0.001 |
| Drug abuse                        | 2 397 (11.2)                                | 26 671 (1.4)      | < 0.001 |
| Hypertension, uncomplicated       | 1 888 (8.8)                                 | 249 682 (13.2)    | < 0.001 |
| Chronic pulmonary disease         | 1 863 (8.7)                                 | 174 216 (9.2)     | 0.01   |
| Congestive heart failure          | 1 431 (6.7)                                 | 176 893 (9.3)     | < 0.001 |
| Atrial fibrillation               | 1 041 (4.9)                                 | 156 032 (8.2)     | < 0.001 |
| Liver disease                     | 1 014 (4.7)                                 | 23 001 (1.2)      | < 0.001 |
| Renal failure                     | 775 (3.6)                                   | 59 475 (3.1)      | < 0.001 |
| Hypertension, complicated         | 433 (2.0)                                   | 33 177 (1.8)      | 0.003  |
| Obesity                           | 380 (1.8)                                   | 28 665 (1.5)      |        |
| Solid tumour without metastasis   | 346 (1.6)                                   | 86 236 (4.6)      | < 0.001 |
| Hypothyroidism                    | 264 (1.2)                                   | 42 992 (2.3)      | < 0.001 |
| Metastatic cancer                 | 254 (1.2)                                   | 63 205 (3.3)      | < 0.001 |
| AIDS/HIV                          | 112 (0.5)                                   | 1 264 (0.1)       | < 0.001 |
| Lymphoma                          | 79 (0.4)                                    | 14 152 (0.8)      | < 0.001 |
| Main diagnosis at index admission |                                             |                  |
| Mental disorder                   | 4 645 (21.7)                                | 102 472 (5.4)     | < 0.001 |
| Injury or poisoning               | 2 862 (13.4)                                | 140 146 (7.4)     |        |
| Complication of pregnancy, childbirth or puerperium | 2 227 (10.4) | 353 508 (18.7) |        |
| Disease of the digestive system   | 2 196 (10.3)                                | 215 699 (11.4)    |        |
| Cardiovascular disorder           | 1 921 (9.0)                                 | 249 211 (13.2)    |        |
| Disease of the respiratory system | 1 445 (6.7)                                 | 137 430 (7.3)     |        |
| Disease of the musculoskeletal system and connective tissue | 607 (2.8) | 91 314 (4.8) |        |
| Neoplasm                          | 358 (1.7)                                   | 118 334 (6.3)     |        |
| Disease of nervous system or sense organs | 332 (1.6) | 41 305 (2.2) |        |
| Infectious or parasitic disease   | 311 (1.5)                                   | 19 007 (1.0)      |        |
| Disease of blood and blood-forming organs | 129 (0.6) | 13 462 (0.7) |        |
| Length of index admission, d, median (IQR) | 3 (2–6) | 5 (3–9) | < 0.001 |
| Deaths within indicated intervals after discharge, d§ | | |
| 7                                 | 97 (0.5)                                    | 6 956 (0.4)       | 0.05   |
| 14                                | 165 (0.8)                                   | 13 145 (0.7)      | 0.2    |
| 30                                | 328 (1.5)                                   | 26 600 (1.4)      | 0.1    |
| 90                                | 739 (3.5)                                   | 67 860 (3.6)      | 0.3    |
| 180                               | 1 167 (5.5)                                 | 111 171 (5.9)     | 0.001  |

Note: AMA = left hospital against medical advice, IQR = interquartile range, non-AMA = did not leave hospital against medical advice, SD = standard deviation. *Unless otherwise indicated. †Only selected variables are shown; see Appendix 2 for full results. §p < 0.05 for tests. $These are not true death rates, because the unit of measure for this analysis was admissions, not patients.
the units of analysis. Cases were all deaths of adults aged 18–90 years that occurred during the study period among people with at least 1 hospital admission from which they were discharged alive in the 90 days before death (the eligibility period). The controls for each case were patients of the same age and sex who were alive on the date of the patient in the case group’s death, and who had at least 1 admission to hospital from which they were discharged alive in the 90 days before that date. We matched each case, without replacement, to up to 5 controls. Patients in each group were required to have survived the recent admission to hospital, because that is a precondition of having left against medical advice. To adjust for confounding in a matched analysis, we used conditional logistic regression, adjusting for year of admission, comorbidities, socioeconomic status, the number of admissions to hospital in the 5 years before the eligibility period and the existence of any AMA discharges in the 5 years before the eligibility period. We excluded age and sex from this regression because cases and controls were matched on those parameters. As a preplanned sensitivity analysis, we included the main hospital diagnosis in the 90-day mortality model. To further explore the association of leaving against medical advice and death, we created similar case–control models for admission to hospital within the following nonoverlapping intervals before death: 0–7, 8–14, 15–30, 31–90 and 91–180 days.

We used χ² tests, t tests, and Mann–Whitney U tests as appropriate for univariate comparisons (SAS Institute, Cary, NC). Our data are presented as mean (± standard deviation), unless otherwise indicated, and we considered p values of less than 0.05 to be statistically significant. Our reported odds ratios (ORs) are adjusted per the multivariable regression models. Analyses were performed with SAS version 9.1.

Results

We identified 1916 104 index admissions to hospital among 610 187 patients, 229 685 of whom (37.6%) had a single admission. Of the total number of admissions we identified, 21 417 (1.1%) ended with patients leaving against medical advice (i.e., an AMA discharge status). Unscheduled readmission within 30 days occurred after 234 809 index admissions (12.3%).

We found several differences between admissions resulting in AMA and non-AMA discharges (Table 1, Appendix 2, available at www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.130029/-/DC1). Without adjusting for these differences, AMA discharge status was associated with double the rate of unscheduled readmission to hospital within 30 days (24.0% v. 12.1%, χ² p < 0.001; Table 2). The difference in readmission rates between admissions ending with AMA and non-AMA discharges grew continuously over time; however, most of the higher readmission rate for AMA discharge status manifested early. Twenty-five percent of the cumulative difference in rates at 180 days after the index admission to hospital occurred within 1 day, and 60% occurred within 2 weeks.

Our logistic regression model shows that AMA discharge status was associated with higher odds of readmission within 30 days (Table 3, Appendix 3, available at www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.130029/-/DC1). The within-person effect was an OR 2.10-fold higher (95% CI 1.99–2.21) when a given patient left against medical advice compared with when he or she did not. The between-person effect shows that for 2 otherwise comparable patients each having a single index admission, the patient who left against medical advice had 3.04-fold higher odds of readmission at 30 days (95% CI 2.79–3.30). Other variables associated with 30-day readmission were older age,}

| Table 2: Rates of unscheduled readmissions to hospital after index admissions from which patients were discharged alive, by discharge status |
|-----------------|-----------------|-----------------|-----------------|
| Time after discharge, d* | AMAn = 21 417 | non-AMAn = 1 894 687 | Difference (95% CI) (AMA v. non-AMA) |
| 1 | 5.5 | 1.2 | 4.3 (4.0–4.6) |
| 7 | 13.6 | 5.1 | 8.5 (8.1–9.0) |
| 14 | 18.1 | 7.9 | 10.2 (9.7–10.7) |
| 30 | 24.0 | 12.1 | 11.9 (11.3–12.4) |
| 60 | 30.5 | 16.9 | 13.6 (13.0–14.3) |
| 90 | 34.9 | 20.1 | 14.8 (14.1–15.4) |
| 180 | 43.3 | 26.4 | 16.9 (16.1–17.5) |

| Interval-specific rates, % |
|-----------------|-----------------|-----------------|-----------------|
| Interval after discharge, d* | AMAn = 21 417 | non-AMAn = 1 894 687 | Ratio (95% CI) (AMA:non-AMA) |
| 0–7 | 5.5 | 1.2 | 4.7 (4.4–5.0) |
| 8–14 | 8.1 | 3.9 | 2.1 (2.0–2.2) |
| 15–30 | 4.5 | 2.8 | 1.6 (1.5–1.7) |
| 31–60 | 5.9 | 4.2 | 1.4 (1.3–1.5) |
| 61–90 | 6.6 | 4.8 | 1.4 (1.3–1.5) |
| 91–181 | 8.4 | 6.3 | 1.3 (1.3–1.5) |

Note: AMA = left hospital against medical advice, CI = confidence interval, non-AMA = did not leave hospital against medical advice. *Includes first and last days of interval.
Table 3: Results of the logistic regression model for unscheduled readmission to hospital within 30 days

| Variable*                                                                 | OR† (95% CI)          |
|---------------------------------------------------------------------------|-----------------------|
| AMA discharge from index admission†                                      |                       |
| Between-person effect                                                    | 3.04 (2.79–3.30)      |
| Within-person effect                                                     | 2.10 (1.99–2.21)      |
| Any AMA discharges in previous 5 yr                                      | 1.25 (1.16–1.34)      |
| No. of admissions in previous 5 yr (per admission)                       | 1.12 (1.11–1.12)      |
| Age (per yr)                                                             | 1.02 (1.02–1.02)      |
| Male sex                                                                 | 1.08 (1.06–1.10)      |
| Household income quintile                                                |                       |
| Urban 1 (lowest urban stratum)                                           | 1.00 (ref)            |
| Urban 3                                                                  | 0.93 (0.90–0.95)      |
| Urban 5 (highest urban stratum)                                          | 0.82 (0.79–0.84)      |
| Rural 1 (lowest rural stratum)                                           | 1.30 (1.26–1.33)      |
| Rural 3                                                                  | 1.26 (1.22–1.29)      |
| Rural 5 (highest rural stratum)                                          | 1.08 (1.05–1.12)      |
| Not calculated                                                           | 0.99 (0.95–1.03)      |
| Comorbid conditions§                                                     |                       |
| Metastatic cancer                                                        | 3.01 (2.91–3.10)      |
| Lymphoma                                                                 | 2.45 (2.31–2.60)      |
| AIDS/HIV                                                                 | 2.41 (1.98–2.94)      |
| Renal failure                                                            | 1.66 (1.60–1.72)      |
| Liver disease                                                            | 1.59 (1.51–1.68)      |
| Solid tumor without metastasis                                           | 1.45 (1.41–1.50)      |
| Congestive heart failure                                                 | 1.44 (1.41–1.47)      |
| Drug abuse                                                               | 1.25 (1.13–1.34)      |
| Diabetes                                                                 | 1.21 (1.19–1.24)      |
| Chronic pulmonary disease                                               | 1.17 (1.15–1.20)      |
| Arrhythmia                                                               | 1.08 (1.06–1.11)      |
| Hypertension, complicated                                               | 1.08 (1.03–1.14)      |
| Hypothyroidism                                                           | 1.05 (1.00–1.10)      |
| Alcohol abuse                                                            | 1.00 (0.95–1.04)      |
| Hypertension, uncomplicated                                              | 0.98 (0.96–0.99)      |
| Obesity                                                                  | 0.95 (0.90–1.00)      |
| Main diagnosis at index admission                                        |                       |
| Cardiovascular disorder                                                 | 1.00 (ref)            |
| Complication of pregnancy, childbirth or puerperium                      | 2.11 (2.05–2.18)      |
| Disease of blood or blood-forming organs                                 | 1.17 (1.04–1.31)      |
| Neoplasm                                                                 | 0.94 (0.91–0.97)      |
| Infectious or parasitic disease                                         | 0.93 (0.89–0.99)      |
| Disease of the digestive system                                         | 0.93 (0.91–0.95)      |
| Mental disorder                                                          | 0.92 (0.89–0.96)      |
| Disease of the respiratory system                                       | 0.90 (0.88–0.92)      |
| Disease of nervous system or sense organs                                | 0.75 (0.70–0.80)      |
| Injury or poisoning                                                      | 0.74 (0.72–0.76)      |
| Disease of the musculoskeletal system or connective tissue              | 0.74 (0.71–0.76)      |
| Length of the index admission (per d)                                   | 0.997 (0.997–0.998)   |

Note: AMA = against medical advice, CI = confidence interval, OR = odds ratio. Comorbid conditions and diagnoses at admission are listed in descending order of OR and, thus, influence on odds of readmission.

*Only selected variables are shown. Please see Appendix 3 for full results.
†Adjusted for between-person and within-person AMA discharge status, age, sex, year of index admission, comorbid conditions, household income, diagnosis of the index admission, length of index admission, number of admissions in the previous 5 yr and existence of any AMA discharges in the previous 5 yr.
‡Within-person effects indicate the difference in readmission risk after an AMA discharge and a non-AMA discharge for the same patient. Between-person effects indicate the difference in readmission risk after an AMA discharge and a non-AMA discharge for one patient and a non-AMA discharge for another patient with otherwise identical characteristics.
§Because a patient can have any number of comorbidities, they are represented in the regression model as individual binary indicator variables, thus no reference condition was chosen.
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disccharges in the previous 5 yr. Exact matching occurred on age and sex.

household income, number of admissions in the previous 5 yr and the existence of any AMA discharges in the previous 5 yr. Exact matching occurred on age and sex.

male sex, a high number of admissions to hospital in the previous 5 years, low socioeconomic status (among both urban and rural residents), short stay in hospital, admission to hospital later in the study period and certain comorbid conditions and hospital diagnoses. The results of the sensitivity analyses were similar to those of our main model (Appendix 4, available at www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.130029/-/DC1).

The case–control analysis of 90-day mortality included 45 848 cases, with at least 1 matched control for 44 837 (97.8%); 60% of cases had 5 controls, for a total of 184 210 controls. As expected, cases and controls differed substantially (Appendix 5, available at www.cmaj.ca/lookup/suppl/doi:10.1503/cmaj.130029/-/DC1). With multivariable adjustment, the odds of death were 2.51-fold higher for patients with AMA discharge status (95% CI 2.16–2.87). Although the odds of death associated with AMA discharge status generally declined over the nonoverlapping intervals after discharge from hospital, it remained significantly elevated even out to 180 days (p < 0.001 for all, Figure 1).

### Interpretation

Leaving the hospital against medical advice was associated with increased risks of readmission to hospital and death that persisted for at least 6 months. Potential mechanisms for these associations directly related to the patients’ acute illness include more severe illness or incomplete treatment of the illness. Such direct effects would be expected to manifest early and decline with time. Other mechanisms could relate to patient characteristics or behaviours that correlate with the tendency to leave hospital against medical advice; such effects could first appear early or later on, but would be expected to persist. The temporal trends we saw allow us to speculate as to these mechanisms. For both hospital readmission and death, the elevated rates among patients who left against medical advice started out high and then declined, but remained elevated to at least 180 days. This trend suggests that mechanisms both directly related to the acute illness (early), and unrelated to the acute illness (persistent) are in effect. The persistent influence may relate to intrinsic, health-related patient characteristics; one possibility is a general tendency to nonadherence with medical recommendations, which has been associated with increased mortality.26–29

Previous studies of the consequences associated with leaving hospital against medical advice, including 2 population-based studies,6,8 were restricted to selected types of patients.28–30 Only 4 of these studies included more than 670 AMA discharges,26–28 whereas we included more than 21 000

### Table 4: Results of the conditional logistic regression model for 90-day mortality

| Variable* | OR† (95% CI) |
|-----------|-------------|
| AMA discharge from index admission | 2.51 (2.18–2.89) |
| Any AMA discharges in previous 5 yr | 0.82 (0.66–1.01) |
| No. of admissions in previous 5 yr (per admission) | 0.69 (0.68–0.69) |
| Household income quintile | |
| Urban 1 (lowest urban stratum) | 1.00 (ref) |
| Urban 2 | 0.95 (0.90–1.00) |
| Urban 3 | 0.89 (0.84–0.94) |
| Urban 4 | 0.85 (0.80–0.90) |
| Urban 5 (highest urban stratum) | 0.81 (0.76–0.87) |
| Rural 1 (lowest rural stratum) | 1.67 (1.58–1.77) |
| Rural 2 | 1.72 (1.63–1.81) |
| Rural 3 | 1.62 (1.53–1.71) |
| Rural 4 | 1.50 (1.41–1.59) |
| Rural 5 (highest rural stratum) | 1.23 (1.15–1.31) |
| Not calculated | 2.24 (2.11–2.38) |

Comorbid conditions

| Metastatic cancer | 27.00 (25.82–28.23) |
| AIDS/HIV | 17.73 (11.77–26.70) |
| Lymphoma | 10.47 (9.52–11.51) |
| Liver disease | 4.51 (4.11–4.96) |
| Solid tumor without metastasis | 4.14 (3.97–4.32) |
| Renal failure | 3.40 (3.19–3.63) |
| Chronic pulmonary disease | 2.26 (2.17–2.35) |
| Drug abuse | 1.89 (1.62–2.20) |
| Diabetes | 1.86 (1.80–1.93) |
| Alcohol abuse | 1.30 (1.20–1.40) |
| Arrhythmia | 1.19 (1.14–1.24) |
| Hypothyroidism | 1.05 (0.98–1.13) |
| Obesity | 0.87 (0.77–0.98) |
| Hypertension, uncomplicated | 0.86 (0.83–0.89) |
| Hypertension, complicated | 0.85 (0.78–0.92) |

Note: AMA = against medical advice, CI = confidence interval, OR = odds ratio.
*Only selected variables are shown; see online Appendix 6 for full results.
†Adjusted for AMA discharge status, year of index admission, comorbid conditions, household income, number of admissions in the previous 5 yr and the existence of any AMA discharges in the previous 5 yr. Exact matching occurred on age and sex.
such events. In addition, most previous studies lacked complete follow-up and determination of outcomes. Three studies evaluated death related to leaving against medical advice. In a study involving hospitals under the direction of the US Department of Veterans Affairs, leaving against medical advice was associated with hazard ratios of 1.1 for 30- and 60-day mortality, but was only significant for the latter period. In data from a single hospital, 6-month mortality for AMA discharges was 19.4%, compared with 3.2% for non-AMA discharges. Lastly, a study involving medical patients in 2 hospitals reported ORs of 2.05–2.46 for 30-day mortality associated with leaving against medical advice, depending on the adjustment method. Similar to our results, 10 previous studies reported that leaving against medical advice was associated with elevated rates of readmission to hospital. However, all but 2 of these studies had unreliable determination of readmission rates because they only included readmission to the index hospitals. In those 2 studies, leaving against medical advice was associated with an OR of 2.7 for 1 month readmission among women in the postpartum period, and an OR of 1.9 among patients admitted for cardiovascular diseases.

The readmission rates we saw are in line with those reported throughout Canada.

**Strengths and limitations**

Our analysis relied on data from a large, population-based sample. Furthermore, we used a validated assessment of leaving against medical advice, engaged in complete follow-up and accounted for data clustering. However, our validation of AMA discharge status in hospital abstracts was based on a small sample, and our identification of AMA status from hospital abstracts was imperfect. As discussed in Appendix 1, consequent misclassifications are not likely to alter our conclusions that leaving against medical advice is associated with harm, because misclassifications generally bias effects toward the null result of no difference between groups.

We were unable to adjust for the severity of acute illnesses because our data contained no such measure. However, this limitation is ameliorated by the overlap in predictive power between measures of acute physiologic derangement and variables for which we did adjust, including comorbidities and acute diagnosis. Furthermore, although we assessed all-cause readmissions and deaths, we do not know how many were for the same reason as the index admission.

Although we have speculated that health behaviours correlated with leaving against medical advice may be causally related to elevated mortality, we did not have data on such characteristics. In addition, ethnicity is associated with leaving against medical advice, but our administrative data do not allow its identification. This is particularly relevant in Manitoba, because the province’s large Aboriginal population is known to have poorer health outcomes than those seen in the general population.

Patients with cancer were overrepresented in our case–control mortality analysis, but this is unlikely to have led to spurious findings because our regression model adjusted for this comorbidity. In addition, any residual bias would tend to be toward the

| Interval after discharge, d | OR (95% CI) |
|-----------------------------|-------------|
| 0–7                         | 2.80 (1.85–4.23) |
| 8–14                        | 2.34 (1.57–3.49) |
| 15–30                       | 2.55 (1.98–3.27) |
| 31–90                       | 2.06 (1.75–2.41) |
| 91–180                      | 1.96 (1.68–2.29) |

Figure 1: Odds of death at given intervals after hospital discharge associated with leaving hospital against medical advice. Values are derived from separate case–control regression analyses. Note: CI = confidence interval, OR = odds ratio.
null, because patients with malignant disease had a low rate of leaving against medical advice (0.3%) and high mortality (OR 4.1–27.0, Table 4).

Finally, our results derive from a Canadian province with a publicly-funded universal health care system and may not be generalizable to other settings.

Conclusion

Leaving hospital against medical advice can have deleterious consequences for patients, including subsequent readmission to hospital and mortality (OR 4.1–27.0, Table 4). Although strategies targeted at trying to convince patients not to leave prematurely might diminish the early effects of leaving against medical advice, reducing the persistently elevated risk will likely require longitudinal interventions extending beyond hospital admission.

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