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Abstract

Lung cancer (LC) mortality, as one of the top cancer deaths in China, has been associated with increased levels of exposure to ambient air pollutants. In this study, different lag times on weekly basis were applied to study the association of air pollutants (PM2.5, PM10, and NO2) and LC mortality in Ningbo, and in subpopulations at different age groups and genders. Furthermore, seasonal variations of pollutant concentrations and meteorological variables (temperature, relative humidity, and wind speed) were analysed. A generalised additive model (GAM) using Poisson regression was employed to estimate the effect of single pollutant model on LC mortality in Yangtze River Delta using Ningbo as a case study. It was reported that there were statistically significant relationships between lung cancer mortality and air pollutants. Increases of 6.2\% (95\% confidence interval).
confidence interval [CI]: 0.2% to 12.6%) and 4.3% (95% CI: 0.1% to 8.5%) weekly total LC mortality with a 3-week lag time were linked to each 10 µg/m³ increase of weekly average PM$_{2.5}$ and PM$_{10}$ respectively. The association of air pollutants (PM$_{2.5}$, PM$_{10}$ and NO$_2$) and LC mortality with a 3-week lag time was also found statistically significant during periods of low temperature ($T < 18^\circ$C), low relative humidity ($H < 73.7\%$) and low wind speed ($u < 2.8$m/s), respectively. The female population was found to be more susceptible to the exposure to air pollution than the male population. In addition, the population with an age of 50 years or above was shown to be more sensitive to ambient air pollutant. These outcomes indicated that increased risk of lung cancer mortality was evidently linked to exposure to ambient air pollutant on a weekly basis. The impact of weekly variation on the LC mortality and air pollutant levels should be considered in air pollution-related health burden analysis.

**Keywords**

Ambient air pollution; Lag time; Lung cancer mortality; Meteorological factors; Modelling; Yangtze River Delta

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1. Introduction

Lung cancer (LC) is characterised by the uncontrolled growth of abnormal cells in the lungs, which may eventually spread to the other organs by a process known as metastasis (Tanoue, 2012). An increase in the risk factor of most LC cases has been shown to be associated with increased levels of tobacco smoking. However, an increased risk of LC has also been associated with exposure to ambient air pollution (Alberg and Samet, 2003). Carcinogens, which
are generated from processes such as industrial combustion and vehicle exhaust, are often transported along with fine particulate matter and dispersed over the environment. These carcinogens are potentially responsible for the increased risk of lung cancer in an urban area with a high population density (Pope et al., 2011; Alberg and Samet, 2003).

According to the Global Burden of Cancer, LC is the top leading cause for cancer mortality in China (Fitzmaurice et al., 2017). As China experiences rapid economic growth over the past few decades, the level of energy consumption, industrial waste, motor vehicle usage and urban development have all increased, which then contribute towards a dramatic increase in ambient air pollution levels. Furthermore, the significant increase of population also requires a higher level of energy consumption where coal is still one of the primary sources of energy utilised by many industries and households. These factors have all contributed towards exacerbating the effects of local air pollution and carbon emission in China (Chen and Xu, 2010).

Ambient particulate matter (PM) pollution, which was ranked as the fifth-leading risk factor for disease burden in China, was also the top leading environmental risk factor for the country in 2015 (Forouzanfar et al., 2016). PM is a complex heterogeneous mixture of solid and liquid particles suspended in the air, varying from a few nm to tens of µm in size (Jin et al., 2016). Sources of PM can be identified as primary emission or secondary formation processes resulting from power plant, oil refineries, residential fuel combustion and construction (Jin et al., 2016). Airborne PM has received a lot of attention in the last few decades as they pose a significant risk to human health. For instance, PM$_{10}$ with an aerodynamic diameter of less than 10 µm can enter the lower respiratory system. PM$_{2.5}$, which is smaller than 2.5 µm in aerodynamic diameter, has gradually received increased attention and awareness as it can penetrate into the gas-exchange area of the lung (Brunekreef and Holgate, 2002). In 2012, PM$_{2.5}$ was added into the Chinese...
National Ambient Air Quality Standards (NAAQS) as a new criterion pollutant in order to monitor air quality control in China (Chen et al., 2015; You, 2014).

The association of exposure to ambient PM pollution with respiratory and cardiovascular diseases such as stroke, ischaemic heart disease (IHD), chronic obstructive pulmonary disease (COPD) and LC has been well analysed in previous studies (Forouzanfar et al., 2016; Yang et al., 2013; Burnett et al., 2014). However, previous studies on the association between LC mortality and exposure to ambient air pollution have been limited and results have been inconsistent. For example, it analysed the positive relationship between the current-day concentration of PM$_{2.5}$ pollutant and LC mortality in Chinese cities including Beijing, Chongqing and Guangzhou (Wang et al., 2019). However, another study conducted in Hebei demonstrated that there was no statistically significant association between LC mortality and PM$_{10}$ (Zhu et al., 2019a). Currently, little is known about the effect of seasonal variation and lag time on the LC mortality attributable to ambient air pollution. The effects of SO$_2$, NO$_2$ and PM$_{10}$ on mortality linked to respiratory diseases mortality were previously found to be strongest at lag times of 1 day, 1 day and 2 days, respectively, using the single-day lag model (Zhu et al., 2019a). Seasonality can also impact the acute effects of ambient PM on mortality (Chen et al., 2013). The mortality attributable to PM$_{10}$ was found at its highest level in winter and lowest in summer (Peng et al., 2005). Previous studies have also demonstrated that exposure to NO$_2$ is linked to an increase in the risk of respiratory mortality or morbidity including respiratory inflammation and lung function impairment (Jiang et al., 2019).

The meteorological variables such as temperature, relative humidity, and wind speed play a significant role when defining the association between air pollutant and disease burden. Previous study conducted in Tianjian pointed out a U-shaped relationship between temperature and mortality, which suggested that both extreme low and high ambient temperature increased
the risk of mortality (Guo et al., 2011). PM concentration was shown to be negatively and non-linearly associated with wind speed (Li et al., 2015). Moreover, temperature and relative humidity are often considered as potential confounding factors which need to be adjusted in the stage of model development (Oo et al., 2020; Qiu et al., 2018; Zhu et al., 2019b). Therefore, the effect of meteorological factors (temperature, relative humidity, and wind speed) was added and adjusted in the model development of this study to obtain a more accurate relationship between air pollutant and LC mortality.

Since previous studies on the relationship between LC mortality and PM have been shown to be inconsistent, further investigation is required on the effect of lag time and variation of meteorological conditions on LC mortality attributable to air pollution. Furthermore, most of the previous studies studied the association between LC mortality and air pollutant using lag time effect in a daily basis, which did not consider the potential accumulative effect of air pollutant on LC mortality on a weekly basis. In order to better capture the significant effect of the air pollutant, a new perspective with the different lag time basis is needed to explore when addressing the association between air pollutant and disease burden. Therefore, in this study, the association between air pollutant and LC mortality is characterised on a weekly basis with the consideration of lag time and meteorological changes. Ningbo, a port city located in the Yangtze River Delta region, is employed as a case study with a focus on available pollutant, meteorological and mortality data collected in 2015. The objectives of this study are to determine (1) the association of air pollutants and LC mortality with different lag times on weekly basis; (2) the seasonal variations of air pollutants, meteorological variables (temperature, humidity and wind speed) and LC mortality; and (3) the trends of LC mortality in a subpopulation according to the stratification of age groups and gender.
2. Methodology

2.1 Study area

Located in the southeast region of the Yangtze River Delta (YRD), Ningbo is a coastal city of Zhejiang province that currently consists of six districts, two county-level cities and two counties. Ningbo has four distinct seasons and its climate is mostly affected by subtropical monsoons. The annual average temperature of the city is reported as 16.9 degrees Celsius with the highest monthly average temperature of 28.3°C in July and lowest monthly average temperature of 5.3°C in January. In 2015, the total number of registered people in Ningbo was approximately 5.86 million with 64.6% of them accounting for the urban population. With a total land area of 9816 km², Ningbo had an overall population density of 598 people/km² and a population density of 943 person/km² in its urban districts (Ningbo Statistical Yearbook 2016).

After the open-door policy was announced by the central government in 1979, the manufacturing industries in Ningbo have grown dramatically due to its ideal geographical location of being a port city, which has contributed to its economic development. This has resulted in significant industrialisation and urbanisation taking place in Ningbo (Tang et al., 2015). By the end of 2015, the total GDP in Ningbo accounted for 801.15 billion RMB with 3.6%, 49.0% and 47.4% of the total GDP contributed by the primary, secondary and tertiary industries, respectively (Ningbo Statistical Yearbook 2016). As a result of rapid industrialisation, Ningbo is experiencing increased levels of respiratory related diseases such as lung cancer. In order to carry out a comprehensive study of the relationship between ambient air pollution and LC mortality, Ningbo was chosen as a case study.
2.2 Data collection and handling

Daily LC mortality data of residents living in Ningbo was collected from Ningbo Centre for Disease Control and Prevention (Ningbo CDC) during the year of 2015. The mortality attributable to lung cancer was classified as C33 (malignant neoplasm of trachea) and C34 (malignant neoplasm of bronchus and lung) according to the International Classification of Diseases, Revision 10 (ICD 10). The daily LC mortality derived from the raw data was used to calculate the total number of LC mortality on weekly basis for the study.

Both air pollutant concentration data and meteorological data in an hourly average basis during the year of 2015 were obtained from the Ningbo Environmental Monitoring Centre (Ningbo EMC). Two monitoring stations, which were Ningbo Taigu primary school (29° 860728’ N, 121° 289801’ E) and Sanjiang middle school (29° 887798’ N, 121° 558176’ E), were selected to represent the air pollutant concentration and meteorological data of Ningbo. As the location of two monitoring stations are near to town centre, which is surrounded by commercial activities and residential areas, the pollutant data is acceptable to describe exposures to air pollutants in Ningbo city. The daily average and weekly average values of both air pollutant concentration and meteorological data were derived from raw data collected for further analysis.

In addition, the weekly average values of air pollutant concentration were measured based on the requirement of the Chinese Ambient Air Quality Standard (GB3095-2012), which validates the effectiveness of air pollutant data, using the following criteria: (1) Hourly average concentrations which were a negative value or equal to zero, or entries lacking comprehensive data were eliminated; (2) Each day must have 20 hours or more hourly average concentration data available; (3) Each month must have at least 27 days or more daily average concentration
data available (at least 25 days or more daily average concentration data must be available in February); (4) 324 days or more daily average concentration data must be available per year.

2.3 Statistical analysis and modelling

The demographic characteristics of the study population were derived during the data collection process. Key pollutants which exceeded the minimum concentration that may pose a health risk to the public were selected by following the standard of the Ambient Air Quality Index (AQI). Furthermore, a study of Pearson’s correlation between the air pollutants was conducted in order to determine their collinearity. A time-series single pollutant model was studied by fitting the generalised additive model (GAM) with Poisson regression as shown in equation (1),

$$\ln[E(m_{ij})] = \alpha + \beta z_j + ns(u, df) + ns(T, df) + ns(H, df) + ns(w, df)$$ (1)

where $z_j$ is the weekly average concentration of pollutant $j$; $E(m_{ij})$ is the expected number of weekly total LC mortality with respect to pollutant $j$ and lag time $i = 0, 1, 2, 3, 4$ week; $\alpha$ is the model intercept; $\beta$ is the corresponding regression coefficient; $u, T, H$ are the weekly average values of wind speed, temperature and relative humidity, respectively; $w$ is the sequential number of week in the study year. The function $ns(\cdot)$ represents natural spline function in the regression model, which was used to control the potential confounding factors such as meteorological variable and time effect. $df$ is degree of freedom, and different values for $u, T, H$ and $w$ are set to 5, 9, 5 and 6, respectively. With the smoothing term in the model fitting, the analyses are adjusted for wind speed, temperature, relative humidity and the effect of week.
For the model fitting, the association of the weekly average pollutant concentration and total number of LC mortality on a weekly basis was examined over the study period of 2015 in the Ningbo city. An odd ratio and 95% confidence intervals of LC mortality with different lag times attributable to each 10 µg/m³ increase in the single pollutant model were examined and compared. Furthermore, the performance of models with different lag times were examined by comparing the root mean square error (RMSE) and Akaike information criterion (AIC). The final model for each pollutant was determined based on the minimum RMSE and AIC, and it was then used for further analysis. The sensitivity analyses were conducted in order to test the robustness of the result by altering the \( df \) for wind speed, temperature, relative humidity and week from 4 to 6, 8 to 10, 4 to 6 and 5 to 7, respectively.

For the seasonal analyses, the data was stratified into four distinct seasons: winter (from December to February; week 1 to 9), spring (from March to May; week 10 to 22), summer (from June to August; week 23 to 35) and autumn (from September to November; week 36 to 48). The effect of meteorological variables such as \( u, T, \) and \( H \) on the behaviour of LC mortality with respect to different lag times were also studied in the following analysis. Daily average values of meteorological variables were employed as the standard value to categorise the data into two different groups (high \( u, T, H \) and low \( u, T, H \)) for discussion purposes. For the stratification analyses, the effect of air pollutant on LC mortality was studied in different subpopulations by gender (male and female) and age (under 50, between 50 and 65, and over 65 years old, respectively). All the analyses were conducted using the mixed GAM computation vehicle (mgcv) and splines packages in R statistical software (version 3.6.1). The result was considered as statistically significant when the measured p-value was less than 0.05 (\( P < 0.05 \)).
3. Results

3.1 Description of the study objects and population exposure

During the study period from 1st January 2015 to 31st December 2015, a total number of 3406 lung cancer (LC) deaths was recorded in the Ningbo, including 2474 deaths (72.6%) from male patients and 932 deaths (27.4%) from female patients. The stratification of LC mortality according to gender and age group is shown in Table 1. By the end of 2015, the total deaths and the deaths due to malignant tumour in Ningbo were recorded as 37838 and 12344 cases, respectively. As a result, 9.0% of death population and 27.6% of deaths due to malignant tumour in Ningbo were associated with lung cancer mortality during the study period.

| Characteristics                  | Total     | Male       | Female     |
|----------------------------------|-----------|------------|------------|
| Total population                 | 5865731   | 2915471 (49.7%) | 2950260 (50.3%) |
| Death population                 | 37838     | 21225 (56.1%)   | 16613 (43.9%)   |
| Deaths due to malignant tumour   | 12344     | 8219 (66.6%)    | 4125 (33.4%)    |
| Lung cancer mortality            | 3406      | 2474 (72.6%)    | 932 (27.4%)     |
| Age of death (<50)               | 135 (4.0%) | 91 (67.4%)     | 44 (32.6%)      |
| Age of death (50-69)             | 1520 (44.6%) | 1125 (74.0%)   | 395 (26.0%)     |
| Age of death (>69)               | 1751 (51.4%) | 1258 (71.8%)   | 493 (28.2%)     |

In Table 2, the standard concentrations which contribute towards an individual ambient quality index (IAQI) value of 50 were based on the Technical Regulation in the Ambient Air Quality Index (HJ 633-2012). All pollutant concentrations were reported as the 24-hour average value, except for $O_3$ which is recorded as the average value per hour. Among the air pollutants, the daily
average concentration of air pollutants PM$_{2.5}$, PM$_{10}$ and NO$_2$ exceeded the concentration, which contributes to the IAQI value of 50, as shown in Table 2. An IAQI value of 50 is the minimum threshold of a potential risk of air pollutant to public health, as indicated in “Standard” column of Table 2. Therefore, PM$_{2.5}$, PM$_{10}$ and NO$_2$ were then selected for further investigation on the association with lung cancer mortality using the single pollutant model. Wind speed, temperature and relative humidity had daily mean values of 2.8 m/s, 18.0°C and 73.7% recorded, respectively. These mean values were used as the cut-offs for dividing the data into high and low levels of meteorological conditions in the stratification analyses. Furthermore, the daily average number of LC mortality was recorded as 9.3, with the highest daily LC mortality value recorded of 20.

**Table 2.** Daily average air pollutant concentration data and meteorological data in Ningbo.

| Variables   | Mean   | SD    | Standard | Min   | Lower quartile | Median | Upper quartile | Max    |
|-------------|--------|-------|----------|-------|----------------|--------|----------------|--------|
| PM$_{2.5}$ (µg/m$^3$) | 44.8   | 25.3  | 35.0     | 2.4   | 26.8           | 38.6   | 55.3           | 153.0  |
| PM$_{10}$ (µg/m$^3$)  | 69.1   | 36.1  | 50.0     | 4.1   | 41.9           | 61.2   | 86.3           | 231.0  |
| CO (mg/m$^3$)         | 0.9    | 0.3   | 2.0      | 0.4   | 0.7            | 0.8    | 1.0            | 1.9    |
| NO$_2$ (µg/m$^3$)     | 45.0   | 17.3  | 40.0     | 11.6  | 31.6           | 42.6   | 57.1           | 109.0  |
| SO$_2$ (µg/m$^3$)     | 18.1   | 7.6   | 50.0     | 6.4   | 13.3           | 16.8   | 21.0           | 68.0   |
| O$_3$ (µg/m$^3$)      | 63.4   | 28.3  | 160.0    | 4.2   | 42.2           | 63.6   | 81.7           | 169.4  |
| u (m/s)               | 2.8    | 1.1   | -        | 0.5   | 2.1            | 2.7    | 3.4            | 6.6    |
| T (°C)                | 18.0   | 7.9   | -        | 2.1   | 10.7           | 19.6   | 24.1           | 33.0   |
| H (%)                 | 73.7   | 13.3  | -        | 34.7  | 64.4           | 74.1   | 85.1           | 100.0  |
| LC mortality          | 9.3    | 3.0   | -        | 1.0   | 7.0            | 9.0    | 11.0           | 20.0   |
3.2 Pearson’s correlation coefficient

Pearson’s correlation coefficient between the pollutants was measured in order to investigate their collinearity in the study, as shown in Table 3. The result showed that all the pollutants are highly correlated to each other with a positive relationship, except for O₃. Table 3 also indicates that the collinearities among all the pollutants are likely to exist, which will affect the accuracy of result for a multiple regression model. Therefore, the effect of single pollutant model on LC mortality was investigated in this study. Furthermore, it was found that all the pollutants except O₃ were negatively associated with meteorological variables such as temperature (T), relative humidity (H) and wind speed (u). This implies that the occurrence of higher pollutant concentrations is likely related to the meteorological conditions of lower temperature, lower relative humidity and lower wind speed.

Table 3. Pearson’s correlation coefficient between air pollutants and meteorological variables.

|       | PM₁₀ | CO  | NO₂ | SO₂ | O₃  | T   | H   | u     |
|-------|------|-----|-----|-----|-----|-----|-----|-------|
| PM₂.₅| 0.972| 0.783| 0.767| 0.750| -0.465| -0.682| -0.422| -0.258|
| PM₁₀ | 0.710| 0.796| 0.798| -0.410| -0.644| -0.479| -0.306|
| CO   | 0.746| 0.560| -0.551| -0.478| -0.176| -0.065|
| NO₂  | 0.729| -0.580| -0.622| -0.271| -0.103|
| SO₂  | -0.335| -0.358| -0.429| -0.291|
| O₃   | 0.640| -0.065| 0.115|
| T    | 0.355| 0.152|
| H    | 0.534|
3.3 Association of air pollutants (PM$_{2.5}$, PM$_{10}$ and NO$_2$) and LC mortality with different lag times

Table 4 shows the association of each 10 µg/m$^3$ increase of pollutant concentration with LC mortality with different lag times. With the minimum AIC and residual errors, the single pollutant model with a 3-week lag time was found to be best model for predicting the weekly total LC mortality from the weekly average pollutant concentration (PM$_{2.5}$ and PM$_{10}$) as compared to the models with other lag times. The result of each model performance (AIC and RMSE) was displayed in Supplemental Material (Table S1). An increase of 6.2% (95% CI: 0.2% to 12.6%) weekly total LC mortality with a 3-week lag time was associated with an increase of 10 µg/m$^3$ weekly average PM$_{2.5}$. In addition, the estimated increase of 4.3% (95% CI: 0.1% to 8.5%) weekly total LC mortality with a 3-week lag time was observed by increasing 10 µg/m$^3$ weekly average values of PM$_{10}$. For the 3-week lag time, the p-values of the single pollutant model for PM$_{2.5}$ and PM$_{10}$ were found to be less than 0.05, which were considered as statistically significant outcome. The detail p-values of the model fitting were reported in Supplemental Material (Table S2-S4). The result showed there were almost no statistically significant associations between air pollutants (PM$_{2.5}$, PM$_{10}$ and NO$_2$) and LC mortality with the lag times other than 3 weeks. Furthermore, the sensitivity analyses also showed the change of degree of freedom ($df$) of key parameters such as temperature, relative humidity, wind speed, and week in the model did not significantly affect the result of the association between LC mortality and pollutant concentration. The result of sensitivity analyses was displayed in Supplemental Material (Table S5-S8). As a result, the effects of air pollutant concentration on LC mortality with a 3-week lag time were investigated further in this study.
Table 4. Odd ratios (and 95 CIs) of lung cancer (LC) mortality with different lag weeks, with respect to each 10 µg/m³ increase in the single pollutant model in Ningbo. (* P < 0.05)

| LC mortality with lag time | PM_{2.5}          | PM_{10}         | NO₂             |
|---------------------------|-------------------|-----------------|-----------------|
| No lag time               | 0.966 (0.913-1.022) | 0.975 (0.938-1.013) | 0.996 (0.939-1.058) |
| 1 week                    | 1.027 (0.971-1.087) | 1.014 (0.975-1.054) | 1.002 (0.944-1.064) |
| 2 weeks                   | 0.939 (0.887-0.993) * | 0.962 (0.925-1.000) * | 0.951 (0.895-1.010) |
| 3 weeks                   | 1.062 (1.002-1.126) * | 1.043 (1.001-1.085) * | 1.046 (0.985-1.112) |
| 4 weeks                   | 1.026 (0.970-1.086) | 1.015 (0.976-1.055) | 1.024 (0.965-1.088) |

Figure 1(a) illustrates the association of the weekly average of PM_{2.5} and the weekly total LC mortality cases with a 3-week lag time. Similar trends were observed over time on both PM_{2.5} concentration and LC mortality with a 3-week lag time, especially for the beginning section of each peak. The patterns for air pollutants PM_{2.5}, PM_{10} and NO₂ were also similar to that of LC mortality over the period of study, as shown in Figure 1(a)-(c). The total number of lung cancer mortality cases on a weekly basis remained in the range of 45 to 83 cases per week. However, it was noted that the trends of both air pollutant and LC mortality during the period from week 34 to week 46 were almost identical but slightly shifted, which may suggest the effect of a different lag time during that particular period, as well as the effect of meteorological variables on the association of air pollutants and LC mortality.
Figure 1. Associations of pollutant concentration (a) PM$_{2.5}$, (b) PM$_{10}$, (c) NO$_2$ and LC mortality with a 3-week lag time.
3.4 Seasonal variations of air pollutants (PM\textsubscript{2.5}, PM\textsubscript{10} and NO\textsubscript{2}), meteorological variables (temperature, humidity and wind speed) and LC mortality

Based on Figure 2(a)-(c), it was observed that the concentrations of air pollutants in winter were much higher and fluctuated more than any other seasons. For example, PM\textsubscript{10} reached a high average concentration of over 97.4 µg/m\textsuperscript{3} and the highest weekly concentration of 145.0 µg/m\textsuperscript{3} occurred in winter. The concentrations of air pollutants were relatively lower and fluctuated less in summer. The average concentrations of PM\textsubscript{2.5}, PM\textsubscript{10} and NO\textsubscript{2} during summer were 27.9 µg/m\textsuperscript{3}, 43.1 µg/m\textsuperscript{3} and 31.1 µg/m\textsuperscript{3} respectively, and the values did not exceed the individual standard concentrations contributing to an IAQI value of more than 50. Moreover, it was detected that the pollutant concentrations in both spring and autumn were moderate and closer to their annual average values. The seasonal variation of meteorological variables was notable in this study. Based on Figure 2(g)-(i), temperature, relative humidity and wind speed in Ningbo were at their lowest levels during winter and highest levels during summer, which provides further evidence on their negative correlation with air pollutants (PM\textsubscript{2.5}, PM\textsubscript{10} and NO\textsubscript{2}).

Figure 2(j) illustrates the seasonal profile of LC mortality in Ningbo. It shows that lower numbers of LC mortality were recorded in summer and autumn, however higher numbers were recorded in winter. In addition, it was noted that there was a strong oscillating effect on the LC mortality in winter, as well as on the concentration of air pollutants over the same period, as shown in both Figure 1 and Figure 2(j). The similar oscillating effects might explain the significant association between air pollutants and LC mortality with a 3-week lag time.
(a) PM$_{2.5}$ concentration

PM$_{2.5}$ concentration over the seasons:
- Spring: ~30 $\mu$g/m$^3$
- Summer: ~20 $\mu$g/m$^3$
- Autumn: ~25 $\mu$g/m$^3$
- Winter: ~40 $\mu$g/m$^3$

(b) PM$_{10}$ concentration

PM$_{10}$ concentration over the seasons:
- Spring: ~60 $\mu$g/m$^3$
- Summer: ~50 $\mu$g/m$^3$
- Autumn: ~45 $\mu$g/m$^3$
- Winter: ~80 $\mu$g/m$^3$

(c) NO$_2$ concentration

NO$_2$ concentration over the seasons:
- Spring: ~40 $\mu$g/m$^3$
- Summer: ~50 $\mu$g/m$^3$
- Autumn: ~60 $\mu$g/m$^3$
- Winter: ~70 $\mu$g/m$^3$

(d) CO concentration

CO concentration over the seasons:
- Spring: ~0.7 mg/m$^3$
- Summer: ~0.8 mg/m$^3$
- Autumn: ~0.9 mg/m$^3$
- Winter: ~1.1 mg/m$^3$

(e) SO$_2$ concentration

SO$_2$ concentration over the seasons:
- Spring: ~15 $\mu$g/m$^3$
- Summer: ~20 $\mu$g/m$^3$
- Autumn: ~25 $\mu$g/m$^3$
- Winter: ~30 $\mu$g/m$^3$

(f) O$_3$ concentration

O$_3$ concentration over the seasons:
- Spring: ~50 $\mu$g/m$^3$
- Summer: ~70 $\mu$g/m$^3$
- Autumn: ~90 $\mu$g/m$^3$
- Winter: ~70 $\mu$g/m$^3$
Figure 2. Weekly average values and their 95% confidence intervals (CIs) of pollutant concentrations (a) PM$_{2.5}$, (b) PM$_{10}$, (c) CO, (d) NO$_2$, (e) SO$_2$, (f) O$_3$; meteorological variables (g) temperature, (h) relative humidity, (i) wind speed; and (j) LC mortality in different seasons.

The model fitting in seasonal analyses showed that the associations between air pollutants PM$_{2.5}$, PM$_{10}$ and NO$_2$, and lung cancer mortality with a 3-week lag time were not statistically significant in all the seasons, as showed in Table 5. However, for the effect of meteorological
changes, it was observed that the individual effect was statistically significant on the relationship between weekly average pollutant concentration and weekly total LC mortality with a 3-week lag during periods of low temperature ($T < 18^\circ C$), low relative humidity ($H < 73.7\%$) and low wind speed ($u < 2.8m/s$) recorded, respectively. For each 10 $\mu g/m^3$ increase in weekly average PM$_{2.5}$, PM$_{10}$ and NO$_2$ concentration, it was observed that respective increases of 5.9\% (95\% CI: 2.0\% to 9.9\%), 4.0\% (95\% CI: 1.4\% to 6.7\%) and 6.9\% (95\% CI: 1.5\% to 12.6\%) in weekly total LC mortality cases occurred with a 3-week lag time when the weekly average temperature was less than 18$^\circ C$. During periods of low relative humidity ($H < 73.7\%$), the weekly total LC mortality with a 3-week lag time increased by 6.3\% (95\% CI: -8.4\% to 23.3\%), 4.1\% (95\% CI: -8.1\% to 17.8\%) and 4.6\% (95\% CI: -10.9\% to 22.7\%), respectively, for each 10 $\mu g/m^3$ increase in weekly average PM$_{2.5}$, PM$_{10}$ and NO$_2$ concentration. Each 10 $\mu g/m^3$ increase in weekly average PM$_{2.5}$, PM$_{10}$ and NO$_2$ concentration in the period of low wind speed ($u < 2.8m/s$) caused 6.6\% (95\% CI: 2.1\% to 11.3\%), 4.0\% (95\% CI: 1.3\% to 6.9\%) and 4.8\% (95\% CI: 0.3\% to 9.6\%) increases in weekly total LC mortality cases with a 3-week lag time. In contrast, it was found there were no statistically significant associations between pollutant concentration and LC mortality during periods of high temperature ($T > 18^\circ C$), high relative humidity ($H > 73.7\%$) and high wind speed ($u > 2.8m/s$).

**Table 5.** Odd ratios (and 95 CIs) of lung cancer (LC) mortality with a 3-week lag time, with respect to each 10 $\mu g/m^3$ increase in single pollutant models (PM$_{2.5}$, PM$_{10}$, NO$_2$) with the effect of seasonal variation and meteorological changes in Ningbo. (* P < 0.05)

| LC Mortality | Lag.3w | PM$_{2.5}$ | PM$_{10}$ | NO$_2$ |
|--------------|--------|------------|-----------|--------|
| **Winter**   |        | 1.052 (0.980-1.130) | 1.038 (0.989-1.090) | 1.075 (0.973-1.189) |
| **Spring**   |        | 1.023 (0.862-1.214) | 0.984 (0.871-1.112) | 0.985 (0.842-1.152) |
|                | Summer       | Autumn       | Autumn       |
|----------------|--------------|--------------|--------------|
|                | 0.920 (0.785-1.077) | 0.946 (0.837-1.070) | 0.924 (0.757-1.128) |
|                | 1.052 (0.973-1.137) | 1.040 (0.989-1.094) | 0.996 (0.927-1.071) |
| T < 18°C       | 1.059 (1.020-1.099) * | 1.040 (1.014-1.067) * | 1.069 (1.015-1.126) * |
| T > 18°C       | 1.043 (0.982-1.107) | 1.039 (0.999-1.080) * | 1.016 (0.950-1.086) |
| H < 73.7%      | 1.063 (0.916-1.233) * | 1.041 (0.919-1.178) * | 1.046 (0.891-1.227) * |
| H > 73.7%      | 0.960 (0.828-1.114) | 0.956 (0.845-1.082) | 0.942 (0.803-1.106) |
| u < 2.8m/s     | 1.066 (1.021-1.113) * | 1.040 (1.013-1.069) * | 1.048 (1.003-1.096) * |
| u > 2.8m/s     | 1.052 (0.999-1.109) | 1.045 (1.006-1.085) | 1.029 (0.959-1.103) |

### 3.5 Stratification analyses by gender and age groups

When stratified by gender, statistically significant associations were found between both genders (males and females) and LC mortality with a 3-week lag time. An increased risk of lung cancer mortality in the female population demonstrated a strong association with increased PM$_{2.5}$ and PM$_{10}$ concentrations, as shown in Table 6. For the female population, each 10 µg/m$^3$ increase in weekly average PM$_{2.5}$ and PM$_{10}$ concentration, the weekly total LC mortality cases increased by 6.7% (95% CI: 0.6% to 13.1%) and 5.0% (95% CI: 1.0% to 9.1%) respectively, with a 3-week lag time. Each 10 µg/m$^3$ increase in weekly average PM$_{2.5}$ and PM$_{10}$ concentration caused the 5.6% (95% CI: 1.8% to 9.4%) and 3.5% (95% CI: 1.0% to 6.0%) increase of weekly total LC mortality in male population, respectively.

The result also indicated that the effects of PM$_{2.5}$, PM$_{10}$ and NO$_2$ were both statistically significant and stronger in the elderly population, especially in the population aged 50 years or above. For the population aged between 50 and 69 years, the weekly total LC mortality cases
were increased by 5.6% (95% CI: 0.9% to 10.6%), 5.0% (95% CI: 1.8% to 8.3%) and 7.8% (95% CI: 1.9% to 14.0%), with the increase of each 10 µg/m³ weekly average PM$_{2.5}$, PM$_{10}$ and NO$_2$ concentration, respectively. Furthermore, the increase of 10 µg/m³ weekly average PM$_{2.5}$ and PM$_{10}$ concentration caused the increase of 7.0% (95% CI: 2.6% to 11.6%) and 3.7% (95% CI: 0.8% to 6.6%) weekly total LC mortality with a 3-week lag time in the population aged above 69 years. In contrast, no significant association were detected between air pollutant and LC mortality in the younger age group which is below 50 years.

Table 6. Odd ratios (and 95 CIs) of lung cancer (LC) mortality with a 3-week lag time, with respect to each 10 µg/m³ increase in single pollutant models (PM$_{2.5}$, PM$_{10}$, NO$_2$) with different gender and age group in Ningbo. (* P < 0.05)

| LC Mortality | Lag.3w |
|--------------|--------|
|               | PM$_{2.5}$ | PM$_{10}$ | NO$_2$ |
| Male         | 1.056 (1.018-1.094) * | 1.035 (1.010-1.060) * | 1.034 (0.990-1.079) |
| Female       | 1.067 (1.006-1.131) * | 1.050 (1.018-1.091) * | 1.058 (0.986-1.136) |
| Age < 50     | 0.944 (0.811-1.099)  | 0.953 (0.861-1.055)  | 0.926 (0.772-1.111) |
| Age 50-69    | 1.056 (1.009-1.106) * | 1.050 (1.018-1.083) * | 1.078 (1.019-1.140) * |
| Age > 69     | 1.070 (1.026-1.116) * | 1.037 (1.008-1.066) * | 1.020 (0.970-1.074) |

4. Discussion

In this study, a single pollutant model with Poisson regression was used to study the association between weekly total lung cancer (LC) mortality cases and weekly average pollutant concentrations such as PM$_{2.5}$, PM$_{10}$ and NO$_2$ in Ningbo, China. With the smoothing function in the model fitting, the effect of LC mortality was adjusted for meteorological factors such as wind speed,
temperature, relative humidity and the effect of week. It was observed that the strongest effects of air pollutants on LC mortality were significant with a lag time of 3 weeks as it resulted in lowest residual errors and AIC values as compared to other models. The risk of weekly total LC mortality with a 3-week lag time was increased by 6.2% and 4.3% for each 10 µg/m³ increase in weekly average PM$_{2.5}$ and PM$_{10}$ concentration, respectively.

Previous studies have shown the important connection between air pollutants and LC mortality. Air pollution exposure has been linked to the generation of reactive oxygen species (ROS) and oxidative damage to DNA that might be associated with the increased risk of lung cancer (Moller et al., 2008). A study using weight-of-evidence approach concluded that exposure to particulate matter may induce direct DNA damage that results in the development of lung cancer (Lynch et al., 2016). To the best of our knowledge, this is the first epidemiological study carried out in China which employs the weekly average pollutant concentrations to evaluate the effect of lag time, meteorological factors and seasonal variation on LC mortality on a weekly basis. Weekly average values of pollutant concentrations were used in this study as the weekly cycle in pollutant concentration was evident in Ningbo. Figure 3(a)-(c) illustrate the weekly cycles of PM$_{2.5}$, PM$_{10}$ and NO$_2$, respectively. Behaviours such as the increased usage of motor vehicles and commercial activities, which contribute towards high pollutant concentration levels, occur in urban areas during Fridays and Saturdays. Therefore, by employing the values on weekly basis, the possible confounding effect caused by the weekly cycle in pollutant concentrations can be minimised during the analyses.
Figure 3. Weekly levels of (a) PM$_{2.5}$, (b) PM$_{10}$ and (c) NO$_2$.

In this study, the increase of 6.2% and 4.3% weekly total LC mortality cases with a 3-week lag time were associated with an increase per 10 µg/m$^3$ weekly average of PM$_{2.5}$ and PM$_{10}$, respectively. The current study was compared with the previous studies, suggesting that there are significantly positive associations between exposure to air pollutants and LC mortality, as shown in Table 7. Previous study showed that the effects of PM$_{2.5}$ and PM$_{10}$ were much stronger in both Guangzhou and Chongqing during their cold season (Wang et al., 2019). Our result was in line with the findings in the previous study, as similar daily averages of pollutant concentrations can be observed in these cities during their study periods. In periods of low temperature, the effect
of particulate matter on LC mortality was much greater than periods of high temperature, indicating that there was strong relationship between air pollutant and temperature. Another study conducted in Shenyang concluded that the effects of PM$_{2.5}$ and SO$_2$ were associated with increased LC mortality levels, especially in the male population, and that exposure to particulate matter could greatly affect the younger population (Xue et al., 2018). Our result showed there is a slightly stronger association between LC mortality and air pollutants in the female population than in the male population, which was different with the previous study (Xue et al., 2018). However, our result was consistent with the findings that female population was found to be more sensitive to exposure to air pollutants than male population (Zhu et al., 2017; Zhu et al., 2019a). A past research conducted in Italy also studied the association between exposure to particulate matter and lung cancer mortality in female population (Uccelli et al., 2016). Another notable result was that exposure to air pollutants significantly affected the elderly population in both age groups of 50-69 years and above 69 years. This is mainly because the efficacy of their immune system levels declines due to aging. Therefore, the elderly population becomes more vulnerable to air pollution-induced health effects (Shumake et al., 2013).

Table 7. Comparison of current study with previous studies conducted in China.

| Author | Pollutant | Increase of daily LCM per 10 µg/m$^3$ increase of daily average pollutant | Lag time | Location | Season |
|--------|-----------|--------------------------------------------------------------------------------|----------|----------|--------|
| Wang et al. (2019) | PM$_{2.5}$ | 0.52% (0.06%-0.99%) | 0-day lag | Beijing, Chongqing, Guangzhou | All year |
| Pollutant | Increase of weekly LCM | Lag time | Location | Season |
|-----------|------------------------|----------|----------|--------|
| PM<sub>2.5</sub> | 6.2% (0.2%-12.6%) | 3-week lag | Ningbo | All year |
| PM<sub>10</sub> | 4.3% (0.1%-8.5%) | 3-week lag | Ningbo | All year |

The seasonal variation of air pollutants was found no statistically significant in this study for PM<sub>2.5</sub>, PM<sub>10</sub> and NO<sub>2</sub>, however it demonstrated peak concentrations during the winter and relatively low concentrations during the summer (Figure 1 and Figure 2). The effect of meteorological changes showed statistically significant outcomes for the association between LC mortality with a 3-week lag time and all pollutants (PM<sub>2.5</sub>, PM<sub>10</sub> and NO<sub>2</sub>, respectively). This may imply that the meteorological variables such as wind speed and relative humidity, which were not consistent within a season, could better explain the relationship between LC mortality and pollutant concentration under different scenarios of meteorological changes. Figure 4 illustrates the profiles of meteorological variables and PM<sub>2.5</sub> concentration in Ningbo. From the perspective of meteorological variables, low humidity and low wind speed levels may contribute towards a poor dispersion environment for pollutants. Therefore, the air pollutants tend to disperse slowly...
and stay for a longer time in the region, which increase the exposure time for its residents. The combined effects of low humidity and low wind speed may potentially exacerbate the influence of air pollutants on LC mortality, especially in winter and autumn where their temperatures are even lower. As shown in Table 5, this study reported that the association between air pollutants and LC mortality was stronger during periods of low temperature, low relative humidity and low wind speed respectively.

Figure 4. Profiles of (a) PM$_{2.5}$ concentration and temperature; (b) Relative humidity and wind speed in Ningbo in 2015.
Conversely, it was noted that the number of LC mortality cases did not reduce in summer when the pollutant concentrations for PM$_{2.5}$, PM$_{10}$ and NO$_2$ were significantly decreased. The average pollutant concentrations were below the IAQI values of 50, which implies almost no significant impact of PM$_{2.5}$, PM$_{10}$ and NO$_2$ on LC mortality during summer. This may be linked to the effect of O$_3$, which generally has a relatively higher concentration in summer due to the higher temperatures and increased amount of sunlight enhancing the production of O$_3$ (Wang et al., 2017). A positive correlation was also observed between O$_3$ and temperature, as shown in Table 3. Ozone is a secondary pollutant and its actual mechanism for affecting human health remains inconclusive. Future works will be required to investigate the individual effect of O$_3$ on the LC mortality, especially during the summer period.

Our model has evidently showed the association between air pollutants (PM$_{2.5}$, PM$_{10}$ and NO$_2$) and LC mortality with the effect of lag time on a weekly basis. Our result may provide some insights into the possible mechanism of air pollutant exacerbating the health condition of a lung cancer patient, which is related to weekly variation of pollutant. Therefore, in order to minimise the impact of air pollution on health burden of the local population, it requires intermediate actions for the air quality improvement such as implementation of stringent regulations and appropriate measure of controlling the pollution from road transport sector especially in urban areas (Angelevska et al., 2021). The introduction of advanced technologies for pollutant removal also plays an important role in air pollution control such as an improved design of catalytic converter for ozone removal (Wu et al., 2019).

The current study has some limitations, as well as recommendation of future works. Firstly, pollutant data and meteorological data were only obtained from two monitoring stations to
represent the actual exposure of the study population, which might lead to measurement errors. Spatial modelling is recommended to estimate the dispersion of pollutant across the city and to improve the accuracy in predicting the impact of air pollutants on LC mortality. Secondly, the effects of indoor air pollution and the smoking habit of patients were not considered due to a lack of relevant data. Previous study showed that the respiratory symptoms and impaired lung functions in female population was linked to higher level of indoor air pollution (Mulenga and Siziya, 2019). It might be a potential confounding factor of the stronger association between LC mortality in female population and air pollutant in this study. Thirdly, a study of lung cancer morbidity was not included, which may underestimate the actual association of air pollution and lung cancer disease. It is suggested that future works employ hospital admission data related to respiratory diseases such as asthma and pneumonia when assessing the impact of air pollution-associated health effects. Finally, the analysis was only performed on Ningbo in 2015 as a case study. Different locations and more study periods should ideally be explored and analysed in order to improve the generalisability of this study.

5. Conclusions

In summary, by using Ningbo as a case study to represent the Yangtze River Delta region, this study has demonstrated and concluded with some important findings below:

(1) It has found statistically significant associations between air pollutants (PM$_{2.5}$ and PM$_{10}$) and lung cancer mortality with a 3-week lag time.

(2) It was observed that high pollutant concentrations were generally found during the winter season, conversely, relatively low pollutant concentrations were found during the summer season.
(3) This study has provided further evidence on the statistically significant association between lung cancer mortality and air pollution during periods of low temperature, low humidity and low wind speed.

(4) The effect of air pollutants on lung cancer mortality was slightly stronger in female population than in male population.

(5) It was observed that the elderly population was more susceptible to exposure to air pollutants, especially those aged 50 years or above with the exposure to particulate matter (PM$_{2.5}$ and PM$_{10}$). It is suggested that more attentions should be given to elderly population and lung cancer patients especially during cold season with low humidity and low wind speed. In addition, more specific mitigation measures against ambient air pollution in terms of advanced technologies and stringent regulations should be introduced and implemented in order to greatly reduce the impact of air pollution on the health burden in China.

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Abbreviations

AIC      Akaike information criterion  
AQI      Air Quality Index  
CI       Confidence interval  
CO       Carbon monoxide  
GDP      Gross domestic product  
IAQI     Individual Air Quality Index  
LC       Lung cancer  
LCM      Lung cancer mortality  
NO$_2$   Nitrogen dioxide  
O$_3$    Ozone  
PM       Particle matter  
PM$_{2.5}$ Particles with an aerodynamic diameter of less than 2.5 µm  
PM$_{10}$ Particles with an aerodynamic diameter of less than 10 µm  
RMSE     Root mean square error  
SD       Standard deviation  
SO$_2$   Sulphur dioxide  

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