A REFINED ANALYSIS OF THE GAP BETWEEN EXPECTED RATE FOR PARTIAL CSIT AND THE MASSIVE MIMO RATE LIMIT

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ABSTRACT

Optimal BeamFormers (BFs) that maximize the Weighted Sum Rate (WSR) for a Multiple-Input Multiple-Output (MIMO) interference broadcast channel (IBC) remains an important research area. Under practical scenarios, the problem is compounded by the fact that only partial channel state information at the transmitter (CSIT) is available. Hence, a typical choice of the optimization metric is the Expected Weighted Sum Rate (EWSR). However, the presence of the expectation operator makes the optimization a daunting task. On the other hand, for the particular, but significant, special case of massive MIMO (MaMIMO), the EWSR converges to Expected Signal covariance Expected Interference covariance based WSR (ESEI-WSR) and this metric is more amenable to optimization. Recently, [1] considered a multi-user Multiple-Input Single-Output (MISO) scenario and proposed approximating the EWSR by ESEI-WSR. They then derived a constant bound for this approximation. This paper performs a refined analysis of the gap between EWSR and ESEI-WSR criteria for finite antenna dimensions.

Index Terms— Beamforming, partial CSIT, EWSR, ESEI-WSR, MaMIMO

1. INTRODUCTION

Interference is the main limiting factor in wireless transmission. Base stations (BSs) with multiple antennas are able to serve multiple Mobile Terminals (MTs) simultaneously, which is called Spatial Division Multiple Access (SDMA) or Multi-User (MU) MIMO. We are particularly concerned here with maximum Weighted Sum Rate (WSR) designs accounting for finite SNR. Typical approaches for maximizing WSR are based on a link to Weighted Sum MSE (WSMSE) [2] or an approach based on Difference of Convex function programming [3] (which is actually better interpreted as an instance of majorization). However, these approaches rely on perfect channel CSIT, which is not practical. Hence, an alternative approach is to maximize the EWSR for the case of partial CSIT.

Partial CSIT formulations can typically be categorized as either bounded error/ worst case (relevant for quantization error in digital feedback) or Gaussian error (relevant for analog feedback, prediction error, second-order statistics information etc.). The Gaussian CSIT formulation with mean and covariance information was first introduced for SDMA (a Direction of Arrival (DoA) based historical precedent of MU MIMO), in which the channel outer product was typically replaced by the transmit side channel correlation matrix, and worked out in more detail for single user (SU) MIMO, e.g. [4]. The use of covariance CSIT was made in the context of Massive MIMO [5], where a not so rich propagation environment leads to subspaces (slow CSIT) for the channel vectors so that the fast CSIT can be reduced to the smaller dimension of the subspace. Such CSIT (feedback) reduction is especially crucial for Massive MIMO. Due to the difficulty in directly optimizing the EWSR metric, optimization of the expected WSMSE (EWSMSE), which is a lower bound for the EWSR, was proposed in [6]. In fact, exact expressions exist for a number of MISO [7] and MIMO cases [8]. However, those expressions are very hard to interpret and to optimize with respect to BFs. This issue has led to the development of large system analysis to try to get simpler expressions for the expected rate [9], [10]. Recently, though under a single user MIMO setting, the authors [11] used a large system approximation for the optimization of the EWSR metric under partial CSIT to counter the impact of Doppler created Inter Carrier Interference (ICI). On the other hand, for the particular, but significant, special case of MaMIMO where the number of transmit antennas is large compared to the number of receive antennas, the EWSR converges to ESEI-WSR and this metric is more amenable to optimization. In another recent publication, [1] considered a multi-user Multiple-Input Single-Output (MISO) scenario and proposed approximating the EWSR by ESEI-WSR. They then derived a constant bound for this approximation. The approximate metric was then used for optimization of the EWSR. Inspired by this, we perform a refined analysis of the gap between EWSR and ESEI-WSR criteria for finite antenna dimensions to evaluate the usefulness of using the ESEI-WSR metric (that is more mathematically tractable) instead of the EWSR.

The main goal of this paper is to show that the much simpler expressions obtained in the ESEI approximation (MaMIMO limit) in fact exhibit only a finite and even small gap to the exact expected rate. Towards this end, we first show in section 2.1 for a general non-zero mean correlated MIMO scenario that the gap is monotonic increasing as a function of SNR and hence is maximum at infinite SNR. Then, we go about deriving this gap at infinite SNR for specific scenarios like uncorrelated MISO (section 3.3), correlated MISO (section 3.4) and uncorrelated MIMO (section 3.5). The swift reduction in the gap with increasing number of antennas is clearly seen for the MISO scenarios. The second order Taylor Series Expansion of EWSR for a general MIMO setting is also derived in section 3.2 and observed to concur with the infinite SNR limits for the gap derived independently. Henceforth, the term gap would refer to the gap between ESEI-WSR and the EWSR. In the following text, the notation $|A|$ refers to the determinant of the matrix $A$. $CN(\mu, C)$ refers to a complex Gaussian distribution with mean $\mu$ and covariance $C$. In this paper, Tx may denote transmit/transmitter/transmission and Rx may denote receive/receiver/reception.

2. MIMO IBC SIGNAL MODEL

Consider an IBC with $C$ cells with a total of $K$ users with $d_k$ streams per user. We shall consider a system-wide numbering of the users. User $k$ has $N_k$ antennas is served by BS $b_k$. The $N_k \times d_k$ received
signal at user $k$ in cell $b_k$ is,

$$y_k = \underbrace{H_{k,b_k} G_k x_k + \sum_{i \neq k, b_i = b_k} H_{k,b_i} G_i x_i + \sum_{j \neq b_k, i = j} H_{k,j} G_i x_i + v_k}_{\text{intracell interf.}} + \underbrace{\sum_{i \neq k} H_{k,b_i} G_i x_i}_{\text{intercell interf.}}$$

(1)

where $x_k$ is the intended (white, unit variance) signal, $H_{k,b_k}$ is the $N_k \times M_{b_k}$ channel from BS $b_k$ to user $k$. BS $b_k$ serves $K_{b_k} = \sum_{i \neq b_k} 1$ users. We consider a noise whitened signal representation so that we get for the noise $v_k \sim \mathcal{CN}(0, I_{N_k})$. The $M_{b_k} \times d_k$ spatial Tx filter or beamformer (BF) is $G_k$.

The scenario of interest is that of partial CSIT available globally with all the BSs. The Gaussian CSIT model for the partial CSIT is

$$H_{k,b_k} = \mathbf{P}_{k,b_k} + \overline{H}_{k,b_k} \mathbf{C}_t^{1/2}$$

(2)

where $\mathbf{P}_{k,b_k} = \mathbf{E}H_{k,b_k}$, and $\mathbf{C}_t^{1/2}$ is the Hermitian square-roots of the Tx side covariance matrices. The elements of $\mathbf{H}_{k,b_k}$ are i.i.d. $\sim \mathcal{CN}(0,1)$.

$$E_{H_{k,b_k}} = (H_{k,b_k} - \mathbf{P}_{k,b_k})(H_{k,b_k} - \mathbf{P}_{k,b_k})^H = \text{tr}(C_i) I_{N_k}$$

(3)

$$E_{\overline{H}_{k,b_k}} = (\overline{H}_{k,b_k} - \mathbf{P}_{k,b_k})(\overline{H}_{k,b_k} - \mathbf{P}_{k,b_k})^H = N_k C_i$$

(4)

Note that the expectation is done over $H_{k,b_k}$, for a known $\overline{H}_{k,b_k}$. This is true for all the expectation operations done in this paper. However, as the parameter over which the expectation is done is clear from the context, henceforth, we just mention the expectation operator $E$ to reduce notational overhead. It is also of interest to consider the total Tx side correlation matrix,

$$EH_{k,b_k}H_{k,b_k}^H = \overline{H}_{k,b_k} \overline{H}_{k,b_k}^H + \text{tr}(C_i) I_{N_k}$$

(5)

2.1. Expected WSR (EWSR)

Once the CSIT is imperfect, various optimization criteria could be considered, such as outage capacity. Here we shall consider the EWSR for a known channel mean $\overline{H}$.

$$EWSR(G) = \mathbf{E} \sum_{k=1}^{K} u_k \ln |1 + G_k^H H_{k,b_k} R_k^{-1} H_{k,b_k} G_k|$$

(6)

$$= \mathbf{E} \sum_{k=1}^{K} u_k \left( \ln |R_k| - \ln |R_k^c| \right) .$$

Here, $G$ represents the collection of BFs $G_k$, $u_k$ are rate weights.

$$R_k = H_{k,b_k} Q_k H_{k,b_k}^H + \mathbf{R}_\pi,$$  
$$Q_k = G_k G_k^H,$$

$$\mathbf{R}_\pi = \sum_{i \neq k} H_{k,b_i} Q_i H_{k,b_i}^H + I_{N_k} .$$

(7)

The EWSR cost function needs to be augmented with the power constraints $\sum_{b_i = j} \text{tr}(Q_{b_i}) \leq P_j$.

2.2. MaMIMO limit and ESEI-WSR

If the number of Tx antennas $M$ becomes very large, we get a convergence for any quadratic term of the form

$$\text{E}QH^H \overset{M \rightarrow \infty}{\longrightarrow} \text{E}QH^H = \overline{Q}H\overline{H}^H + \text{tr}(C_i) I$$

and hence we get the following MaMIMO limit matrices

$$\mathbf{R}_k = \mathbf{R}_\pi + \sum_{k,b_k} \overline{H}_{k,b_k} \overline{H}_{k,b_k}^H + \text{tr}(Q_k C_t k, b_k) I_{N_k}$$

$$\mathbf{R}_\pi = I_{N_k} + \sum_{i \neq k} \left( \overline{H}_{k,b_i} \overline{H}_{k,b_i}^H + \text{tr}(Q_i C_t k, b_i) I_{N_k} \right)$$

(8)

Now, typical approaches to solve the WSR (e.g. the DC approach in [3]) can be run to obtain the max EWSR BF. We shall refer to this approach as the ESEI-WSR approach as (channel dependent) signal and interference covariance matrices are replaced by their expected values. In the following sections, we analyze the gap between the EWSR and the ESEI-WSR to suggest an approximation of the first by the latter in the design of the BF. We would like to remark here that the ESEI-WSR may also be interpreted as the WSR that would be obtained if we assume that the received signal and interference are also Gaussian.

3. EWSR to ESEI-WSR GAP ANALYSIS

We are interested in bounding the difference between ESEI-WSR and the EWSR. At the level of each user $k$, we stack the channel estimates relevant for each user $k$.

$$\mathbf{H}_k = [H_{k,b_1} \cdots H_{k,b_{k-1}} H_{k,b_k} H_{k,b_{k+1}} \cdots H_{k,b_K}]$$

(9)

$$= \mathbf{P}_k + \overline{H}_k \mathbf{C}_t^{1/2}_k,$$

where the elements of $\overline{H}_k$ are i.i.d. $\sim \mathcal{CN}(0,1)$ and $\mathbf{P}_k$ refers to the mean part of $H_{k,b}$. $C_t^{1/2}_k$ is a block diagonal matrix whose $i^{th}$ diagonal block is $C_{t,k,b_i}$. Let $Q_k$ be a block diagonal matrix with each diagonal block being $Q_{b_i}$. $Q_{b_i}$ is similar to $Q$ but with the $k^{th}$ block diagonal set to all zeros. Then,

$$\mathbf{R}_k = \mathbf{I} + \mathbf{H}_k Q_k H_k^H, \quad \mathbf{R}_\pi = \mathbf{I} + \mathbf{H}_k Q_k H_k^H$$

(10)

EWSR(G) = $\sum_{k=1}^{K} \mathbf{E}u_k \ln |\mathbf{R}_k| - \ln |\mathbf{R}_\pi| \right)

$$= \mathbf{E} \sum_{k=1}^{K} \ln |1 + \mathbf{H}_k Q_k H_k^H| - \ln |1 + \mathbf{H}_k Q_k H_k^H|$$

ESEI-WSR(G) = $\sum_{k=1}^{K} \ln |1 + \mathbf{E} \mathbf{H}_k Q_k H_k^H| - \ln |1 + \mathbf{E} \mathbf{H}_k Q_k H_k^H|$

(11)

(12)

Thus, the EWSR and ESEI-WSR have been rewritten in a convenient format so that one can focus on the gap between the two by comparing terms of the form $\mathbf{E} \ln |1 + \mathbf{H}_k Q_k H_k^H|$ and $\ln |1 + \mathbf{E} \mathbf{H}_k Q_k H_k^H|$.

3.1. Monotonicity of gap with SNR

For an SNR $\rho$, define

$$\Gamma(\rho) = \ln |1 + \rho \mathbf{H}_k^H H_k^H| - \mathbf{E} \ln |1 + \rho \mathbf{H}_k^H H_k^H|$$

where $H_k^H \sim \mathcal{CN}(\overline{H}_k^H, C), \overline{H}_k^H = \frac{1}{\sqrt{\rho}} \overrightarrow{P}_k \mathcal{Q}_{\frac{1}{2}},$ and $C = \frac{1}{\rho} \overrightarrow{C}_t^{\frac{1}{2}} \mathcal{Q} \overrightarrow{C}_t^{\frac{1}{2}}$

Then, $1 + \mathbf{H}_k Q_k H_k^H = 1 + \rho \mathbf{H}_k^H H_k^H$

Theorem 1. $\Gamma(\rho)$ is monotonically increasing in $\rho$.
Proof. By Jensen’s inequality, \( \Gamma_k \geq 0 \). To show the monotonicity, we show that the derivative with respect to \( \rho \) is always non-negative. We omit the subscripts and superscripts on \( \mathbf{H} \) for convenience.

\[
\frac{\partial}{\partial \rho} \left( \ln |\mathbf{I} + \rho \mathbf{EHH}^H| - \mathbf{E} \ln |\mathbf{I} + \rho \mathbf{HH}^H| \right) = \\
\text{tr} \left( (\mathbf{I} + \rho \mathbf{EHH}^H)^{-1} \mathbf{EHH}^H - \mathbf{E} \left( (\mathbf{I} + \rho \mathbf{HH}^H)^{-1} \mathbf{HH}^H \right) \right)
\]

(14)

Noting that, \( (\mathbf{I} + \rho \mathbf{EHH}^H)^{-1} \mathbf{EHH}^H \) can be written as \( \frac{1}{\rho} \mathbf{I} - \frac{1}{\rho^2} (\mathbf{I} + \rho \mathbf{EHH}^H)^{-1} \),

\[
\frac{\partial}{\partial \rho} \left( \ln |\mathbf{I} + \rho \mathbf{EHH}^H| - \mathbf{E} \ln |\mathbf{I} + \rho \mathbf{HH}^H| \right) = \\
\frac{1}{\rho} \mathbf{E} \left( (\mathbf{I} + \rho \mathbf{HH}^H)^{-1} \right) - \frac{1}{\rho^2} (\mathbf{I} + \rho \mathbf{EHH}^H)^{-1} \geq 0
\]

(15)

where we have applied Jensen’s inequality as \( (\mathbf{I} + \rho \mathbf{HH}^H)^{-1} \) is a convex function.

As a result, the largest value of \( \Gamma \) will be observed at infinite SNR for a general non-zero mean MIMO with channel \( \mathbf{H} \) with arbitrary transmit covariance matrix. Now, following the same steps as in (11), we can obtain,

\[
\text{ESEI-WSR} - \sum_{k=1}^{K} u_k \Gamma_k(\infty) \leq \text{ESEI-WSR} - \sum_{k=1}^{K} u_k \Gamma_k(\rho) \\
\leq \text{EWSR} \leq \\
\text{ESEI-WSR} + \sum_{k=1}^{K} u_k \Gamma_k(\rho) \leq \text{ESEI-WSR} + \sum_{k=1}^{K} u_k \Gamma_k(\infty).
\]

(16)

In the above, \( \Gamma_k \) and \( \Gamma_k(\infty) \) are terms corresponding to the first and the second terms of equation (11). Remains now to obtain the \( \Gamma(\infty) \) for different scenarios. However, we first look at the Taylor series expansion of EWSR to gain further insight.

3.2. Second-Order Taylor Series Expansion of EWSR

Consider the Taylor series expansion for matrices \( \mathbf{X}, \mathbf{Y} \) of dimension \( N_k \times M \).

\[
\ln |\mathbf{X} + \mathbf{Y}| \approx \ln |\mathbf{X}| + \text{tr} \mathbf{X}^{-1} \mathbf{Y} - \frac{1}{2} \text{tr} \mathbf{X}^{-1} \mathbf{Y} \mathbf{X}^{-1} \mathbf{Y}
\]

(17)

Consider \( \mathbf{X} + \mathbf{Y} = \mathbf{I} + \rho \mathbf{HH}^H, \mathbf{H} = \mathbf{H} + \mathbf{HC}, \mathbf{H} \sim \mathcal{CN}(0, \mathbf{I}) \). For expansion around \( \mathbf{I} + \rho \mathbf{EHH}^H \), choose \( \mathbf{X} = \mathbf{I} + \rho \mathbf{EHH}^H \), \( \mathbf{Y} = \rho (\mathbf{EHH}^H - \mathbf{HH}^H) \). Hence, we get,

\[
\mathbf{E} \ln |\mathbf{I} + \rho \mathbf{HH}^H| \approx \ln |\mathbf{I} + \rho \mathbf{EHH}^H| - \\
\frac{\rho^2}{2} \mathbf{E} \text{tr} \left( \mathbf{X}^{-1} (\mathbf{HH}^H - \mathbf{EHH}^H) \mathbf{X}^{-1} (\mathbf{HH}^H - \mathbf{EHH}^H) \right)
\]

(18)

Using 4th order Gaussian moments [1], we get

\[
\mathbf{E} \ln |\mathbf{I} + \rho \mathbf{HH}^H| \approx \ln |\mathbf{I} + \rho \mathbf{EHH}^H| - \\
\frac{\rho^2}{2} \left( \text{tr} \left( \mathbf{X}^{-1} \right)^2 \mathbf{C}^2 \right)
+ 2\text{tr} \left( (\mathbf{X}^{-1})^2 \mathbf{H}^H \mathbf{X}^{-1} \mathbf{HC} - (\mathbf{H}^H \mathbf{X}^{-1} \mathbf{H})^2 \right).
\]

(19)

Let us denote this second order approximation by \( \Gamma_2(\rho) \), i.e.,

\[
\Gamma_2(\rho) = \frac{\rho^2}{2} \left( \text{tr} \left( \mathbf{X}^{-1} \right)^2 \mathbf{C}^2 \right)
+ 2\text{tr} \left( (\mathbf{X}^{-1})^2 \mathbf{H}^H \mathbf{X}^{-1} \mathbf{HC} - (\mathbf{H}^H \mathbf{X}^{-1} \mathbf{H})^2 \right).
\]

(20)

3.3. MISO independent and identically distributed (iid) channel

In the MISO iid channel, the relevant metric is of the form \( \ln(1 + ||\mathbf{h}||^2) \), where \( \mathbf{h} \) is the \( 1 \times M \) MISO channel vector.

Theorem 2.

\[
0 \leq \ln(1 + \rho ||\mathbf{h}||^2) - \mathbf{E} \ln(1 + \rho ||\mathbf{h}||^2) \leq \gamma - \left( \sum_{i=1}^{M} \frac{1}{\pi_{ii}} - \ln(M) \right) + \frac{1}{M},
\]

(24)

where \( \rho \) is the SNR, \( \gamma \) is Euler constant.

Proof. The proof is given in Appendix [A] .

Note that for \( M = 1 \), the bound reduces to that in (11), namely \( \gamma \). Thus, this bound is a much more refined and tighter bound than what is provided in [1].

We further explore the bound using the properties of the harmonic series. Define \( \mathcal{H}_M = \sum_{k=1}^{M} \frac{1}{k} \). It is known that,

\[
\mathcal{H}_M = \ln(M) + \gamma + \frac{1}{2M^2} - \frac{1}{12M^2} + \frac{1}{120M^4} + \cdots
\]

(25)

Using this in (22), we get

\[
\gamma - (\mathcal{H}_M - \ln(M)) + \frac{1}{M} = \frac{1}{2M^2} + \frac{1}{12M^2} - \frac{1}{120M^4} + \cdots
\]

(26)

Thus, we see that the second order term for the bound is \( \frac{1}{2M^2} \), which is also in agreement with equation (22). In the iid case, \( \mathcal{C} = 1_M \), hence,

\[
\frac{1}{2} \left( \text{tr} \mathbf{C}^2 \right) = \frac{1}{2} \left( \frac{1}{2} \sum_{i=1}^{M} \frac{1}{i^2} \right) = \frac{1}{2M}
\]

(27)

3.4. MISO correlated channel

Theorem 3.

\[
0 \leq \ln(1 + \rho \sum_{i=1}^{p} \lambda_i) - \mathbf{E} \ln(1 + \rho ||\mathbf{h}||^2) \\
\leq \gamma - \left( \sum_{i=1}^{p} \frac{1}{\pi_{ii}}(1 - \lambda_i/\lambda_{\max}) - \ln(p) \right)
\]

(28)

where \( \rho \) is the SNR, \( \gamma \) is Euler constant, \( \lambda_1, \cdots, \lambda_p \) are the non-zero eigenvalues of the correlation matrix \( \mathbf{E HH}^H \).

Proof. The proof is given in Appendix [A].

\]
From the second order Taylor series expansion (equation (22)), the second order term of this bound is

\[
\frac{1}{2} \text{tr}(C^2) \left( \frac{\text{tr}(C)}{1} \right)^2 = \frac{\sum_{i=1}^{p} \lambda_i^2}{2 \left( \sum_{i=1}^{p} \lambda_i \right)^2}
\]

(29)

3.5. MIMO zero mean i.i.d channel

In a multi-user scenario, the regime of interest is \( M \geq N_k \). To tackle this scenario, we first introduce the LDU (Lower Diagonal Upper triangular factorization) of the channel Gram matrix,

\[
\mathbf{HH} = \mathbf{LDL}^H = (\mathbf{LD})^2(\mathbf{LD})^H
\]

(30)

where \( \mathbf{L} \) has unit diagonal and \( \mathbf{D} \) is a diagonal matrix with diagonal entries \( (\mathbf{D}_i) \) greater than zero. The second factorization corresponds to a Cholesky decomposition. The Cholesky factorization of a Wishart matrix (such as \( \mathbf{HH}^H \)) leads to,

\[
\begin{align*}
\mathbf{D}_i & \sim \chi^2(M - i + 1), i \in 1 \cdots N_k \\
\mathbf{L}_{i,j} \mathbf{D}_i & \sim \mathcal{C}N(0,1), i > j
\end{align*}
\]

(31)

which is also known as Bartlett’s decomposition \([13]\). Note that \( |\mathbf{HH}^H| = |\mathbf{LD}_i^H| = |\mathbf{D}_i| \). Hence, \( \ln |\mathbf{HH}^H| = \sum_{i=1}^{N_k} \ln |\mathbf{D}_i| \) and the MIMO case reduces to a sum of MISO scenarios, each having a \( \chi^2 \) distribution with a reducing number of degrees of freedom. Thus, reusing the results in section \([3]\), we get

\[
\Gamma(\infty) = \sum_{i=1}^{N_k} \left( \sum_{k=1}^{M-i+1} \frac{1}{k} - \ln(M - i + 1) \right) + \frac{1}{M - i + 1} + N_k \ln(M) + \sum_{i=1}^{N_k} \ln(M - i + 1)
\]

(32)

\[
= \sum_{i=1}^{N_k} \left( \gamma - \sum_{k=1}^{M-i} \frac{1}{k} - \ln(M) \right)
\]

where the second term addresses the fact that the MISO gaps in section \([3]\) were computed with respect to the ESEI-WSR limit of \( \ln(1 + \rho(M - i + 1)) \), whereas in the MIMO zero mean i.i.d scenario, the ESEI-WSR limit is \( N_k \ln(1 + \rho M) \). For illustration, let us also consider \( M' \gg N_k \). Then using the approximation of the Harmonic series, it can be easily shown that \( \Gamma(\infty) \approx \frac{\pi^2}{6M^2} \), which concurs with the second order Taylor series term in \([22]\).

4. NUMERICAL RESULTS

Figure 1 verifies the infinite-SNR bounds for MISO i.i.d scenario by comparing them against the true values of the gap for different SNRs and different values of \( M \). The true values of the gap are obtained by explicitly performing the integration in Matlab. As expected, the gap is zero at very low SNR. As the SNR increases, the gap monotonically increases to the infinite SNR limit, as predicted in section \([3]\). In addition, the gap reduces rapidly with increasing \( M \). As the MIMO i.i.d case is a sum of MISO i.i.d scenarios, these curves apply to the MIMO i.i.d scenario as well.

Further, to verify the goodness of the second order Taylor series approximation, Figure 2 compares the true gap to the gap approximated from the Taylor series expansion for a zero mean correlated MIMO scenario. This scenario is chosen as we expect gap to be maximum here. The number of receive antennas for each user was chosen as \( N_k = 4 \). \( \rho \) was chosen as 1000. As expected, the Taylor series approximation becomes more accurate with increasing number of Tx antennas. Indeed, even in this MIMO correlated scenario, the gap also reduces quickly as the number of Tx antennas increase.

5. CONCLUSION

In this paper, we have motivated the use of the ESEI-WSR metric (or the MaMIMO limit of the EWSR) for utility optimization involving partial CSIT. Towards this end, we presented a refined bound for the gap between EWSR and the ESEI-WSR. We first showed that the gap is maximum at infinite SNR. The results clearly show that the gap reduces with the number of transmit antennas - thereby concurring with the well known result for the MaMIMO limit. The general
We note the following, is closely approximated by \( \Gamma(\infty) \) then \( \Gamma(\rho) \) should be closely approximated by \( \Gamma_2(\rho) \) also. We can also observe that whenever the gap \( \Gamma(\rho) \) gets small, the second-order term \( \Gamma_2(\rho) \) becomes good, in the sense that \( \Gamma(\rho) = \Gamma_2(\rho) + O(\Gamma_2^2(\rho)) \).

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**A. COLLECTION OF PROOFS**

*Proof. of Theorem 2* To ease the notation, we take \( x = ||h||^2 \), where \( x \) is Chi-squared distributed with mean \( M \). For a \( \chi^2 \) distribution with mean \( M \) and \( 2N_t \) degrees of freedom,

\[
f_X(x) = \frac{x^{M-1}e^{-x}}{(M-1)!}.
\]

As \( \gamma = -\int_0^\infty e^{-x} \ln(x)dx \), at high SNR (\( \rho \to \infty \)),

\[
\mathbb{E}_x \ln(1 + \rho x) = \int_0^\infty f_X(x) \ln(\rho x)dx = \int_0^\infty \frac{x^{M-1}}{(M-1)!} \ln(x)dx + \ln(\rho).
\]

We note the following,

\[
\int e^{-x} \ln(x)dx = -e^{-x} \ln(x) + \text{Ei}(-x), \text{Ei}(x) = -\int_{-x}^\infty e^{-t} \frac{dt}{t}
\]

\[
-\int_0^\infty \frac{x^{M-2}}{(M-2)!} \text{Ei}(-x)dx = \int_0^\infty \frac{x^{M-2}}{(M-2)!} \int_x^\infty e^{-t} \frac{dt}{t}dt
\]

\[
= \int_0^\infty \left( \int_0^x \frac{t^{M-2}}{(M-2)!} \right) e^{-t} \frac{dt}{t}
\]

\[
= \int_0^\infty \frac{x^{M-1}}{(M-1)!} \frac{1}{M-1} = \frac{1}{M-1}
\]

Integrating by parts \( (M \geq 2) \),

\[
\int_0^\infty \frac{x^{M-1}}{(M-1)!} \ln(x)dx = -\frac{x^{M-1}}{(M-1)!} \ln(x) + \text{Ei}(-x) \bigg|_0^\infty - \int_0^\infty \frac{x^{M-2}}{(M-2)!} \left( -e^{-x} \ln(x) + \text{Ei}(-x) \right)\]

The first part in the above equation is zero, so we only need to focus on the second portion of the integral.

\[
\int_0^\infty \frac{x^{M-1}}{(M-1)!} e^{-x} \ln(x)dx = -\int_0^\infty \frac{x^{M-2}}{(M-2)!} \left( -e^{-x} \ln(x) + \text{Ei}(-x) \right)
\]

\[
= \int_0^\infty \frac{x^{M-2}}{(M-2)!} \ln(x)dx - \int_0^\infty \frac{x^{M-2}}{(M-2)!} \text{Ei}(-x)dx
\]

\[
= \int_0^\infty \frac{x^{M-2}}{(M-2)!} e^{-x} \ln(x)dx + \frac{1}{M-1}
\]

The above is a recursive equation, from where, we quickly deduce that,

\[
\int_0^\infty \frac{x^{M-1}}{(M-1)!} e^{-x} \ln(x)dx = \int_0^\infty e^{-x} \ln(x)dx + \sum_{k=1}^{M-1} \frac{1}{k}
\]

\[
= -\gamma + \sum_{k=1}^{M-1} \frac{1}{k}
\]

Thus, we can now write (34) as,

\[
\mathbb{E}_x \ln(1 + \rho x) = \int_0^\infty \frac{x^{M-1}}{(M-1)!} e^{-x} \ln(x)dx + \ln(\rho)
\]

\[
= -\gamma + \sum_{k=1}^{M-1} \frac{1}{k} + \ln(\rho)
\]

\[
= -\gamma + \left( \sum_{k=1}^{M-1} \frac{1}{k} - \ln(M) \right) + \ln(M\rho)
\]

*Proof. of Theorem 3.*

For a correlated MISO scenario, we can write equivalently,

\[
\ln(1 + \rho ||h||^2) = \ln(1 + \rho \sum_{i=1}^p \lambda_i ||h_i||^2),
\]

where \( \lambda_i, i \in \{1 \ldots p\} \) are the non-zero eigen values of the correlation matrix \( \mathbf{E} \mathbf{h} \mathbf{h}^H \), scaled in such a manner that \( \sum_{i=1}^p \lambda_i = M, \mathbf{h}_i \sim \mathcal{CN}(0, 1) \). We make the reasonable assumption that all the non-zero eigen values are unequal. In this case, the probability distribution is given (14) as \( \sum_{i=1}^p \lambda_i e^{-\frac{\pi_i}{\lambda_i}} \), where \( x = \sum_{i=1}^p \lambda_i ||h_i||^2 \).

Thus, at high SNR (\( \rho \to \infty \)),

\[
\mathbb{E}_x \ln(1 + \rho x) = \int_0^\infty \sum_{i=1}^p \lambda_i e^{-\frac{\pi_i}{\lambda_i}} \ln(x)dx + \ln(\rho)
\]

\[
= \sum_{i=1}^p \lambda_i e^{-\frac{\pi_i}{\lambda_i}} \ln(x)dx + \ln(\rho)
\]

\[
= \sum_{i=1}^p \gamma + \pi_i \ln(1 - \lambda_i) + \ln(\rho)
\]

\[
= -\gamma + \left( \sum_{i=1}^p \pi_i \left( 1 - \frac{\lambda_i}{\lambda_f} \right) \right) + \ln(\rho)
\]

\[
(42)
\]
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