A new nonparametric composite exponentially weighted moving average sign control chart

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Abstract. This study proposed a dual Exponentially Weighted Moving Average (EWMA) statistics based nonparametric sign control chart to monitor the process deviation from the targeted value. The performance of the proposed chart is compared with the existing nonparametric version of the EWMA sign and Cumulative Sum (CUSUM) mean control charts by using out-of-control average run length for various shifts in the process. The simulation study showed the superiority of the suggested chart over the existing counterparts. A real dataset is also considered for the application of the proposed chart.

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1. Introduction

Statistical Process Control (SPC) plays an essential role in monitoring and identifying the changes in the process by using well-established statistical tools. These changes are measured in terms of variations, which can be classified into common and assignable causes of variations. The common causes of variations represent an inherent part of the process, leading to the state of statistical in-control [1]. The SPC has gained a considerable progress after the development of the Shewhart control chart [2]. Subsequently, the Cumulative Sum (CUSUM) control chart [3] and the Exponentially Weighted Moving Average (EWMA) control chart [4] have made a considerable improvement. The Shewhart control chart is used in detecting larger shifts in the process as it is based on current sample information only, whereas the CUSUM and EWMA control charts utilize both the current and previous observations and, therefore, are more sensitive to identifying small to moderate shifts.

In order to combine the advantages of the Shewhart control charts in detecting larger shifts and the CUSUM control charts in detecting smaller shifts, the combined Shewhart-CUSUM chart was developed by adding Shewhart limits in the CUSUM chart [5]. A similar scheme was proposed by Lucas and Sacucci [6] by combining the Shewhart scheme with the EWMA control chart. Moreover, Abbas et al. [7] proposed a mixed EWMA-CUSUM control chart to enhance the sensitivity of the conventional memory control charts for smaller shifts. In the same context, Haq [8,9] developed a Hybrid Exponentially Weighted Moving Average (HEWMA) control chart by using dual EWMA charting statistics. This methodology of
combining charts allows dealing with a range of shift values effectively. To be more precise, these charts detect shifts more quickly than the classical CUSUM, EWMA, and traditional Shewhart control charts.

In the aforementioned control charts, the critical quality characteristic of the process is assumed to be normally distributed. However, in many practical situations, the distribution of the underlying process is either unknown or non-normal. This, as a result, seriously affects the efficiency of the typical control charts [10]. One of the possible alternatives in this scenario is the use of the distribution-free or nonparametric control charts which are generally robust, flexible, and more effective. According to Chakraborti and Graham [11], the In-Control (IC) average run-length of all continuous distributions remains the same in distribution-free or nonparametric control charts.

In the context of nonparametric control charts, a new idea was proposed by Yang and Cheng [12] by introducing the process proportion, and a nonparametric CUSUM sign chart was suggested. Subsequently, Yang et al. [13] designed the nonparametric EWMA sign control chart and indicated that their suggested chart was capable of detecting small shifts in the target process when the underlying distribution is either non-normal or unknown. The adaptive Phase-II nonparametric EWMA control chart was proposed by Liu et al. [14] to detect a range of shifts in the location parameter. Lu [15] developed a nonparametric extension of the EWMA control chart called Generally Weighted Moving Average (GWMA) sign control chart to enhance the detection ability for small process shifts. Abbasi et al. [16] investigated the performance of the nonparametric EWMA control chart by using asymptotic, time-varying, and fast initial response based control limits. A mixed version of the EWMA-CUSUM sign chart was proposed by Abbasi et al. [17] using arcsine transformation and they demonstrated that the performance of their chart outperformed other counterparts in terms of shift detection. The synthetic version of the EWMA sign control chart was proposed by Haq [18] by integrating the existing EWMA sign control chart with the conforming run length. Recently, Raza et al. [19] proposed the nonparametric homogeneously weighted moving average control charts based on sign and signed-rank statistics. For more details of nonparametric control charts, we refer the interested readers to, for example, Altu˘lug˘e [20,21], Amin et al. [22], Asham et al. [23], Bekir [24–26], Chowdhury et al. [27], Graham et al. [28–31], Kontras and Triantafyllopoulou [32], Mahmood et al. [33], Raza et al. [34], and the references cited therein.

In this paper, a nonparametric Composite Exponentially Weighted Moving Average (CEWMA) control chart was proposed by combining two EWMA statistics that enhance the sensitivity of the chart. Note that Haq [8,9] designed an HEWMA control chart using two EWMA statistics under the assumption that the underlying distribution of the quality characteristic is normal. However, this chart is not effective in case of violation of normality assumption or when the distribution is unknown which happens to be the most practical situation. Furthermore, the variance expression derived by Haq [8,9] becomes undefined when $\lambda_1 = \lambda_2$. The nonparametric version of a double EWMA control chart was proposed by Riaz and Abbasi [35] using arcsine transformation, and the same value of smoothing constant $\lambda$ was assumed for both EWMA statistics. This makes the use of their chart limited for practitioners if they want to use different values for smoothing parameters. On the contrary, our proposed chart is simple as it considers a sign test without any transformation to define a composite EWMA charting statistic. Furthermore, the CEWMA chart provides greater flexibility than the one proposed by Riaz and Abbasi [35] by allowing practitioners to consider all possible combinations of smoothing parameters $\lambda_1$ and $\lambda_2$. Besides, we have evaluated the performance of the proposed control chart under various distributions, which had not been considered by Riaz and Abbasi [35].

The rest of this paper is organized as follows. The design structure of the proposed CEWMA sign control chart is given in Section 2. Performance evaluations of the proposed charts are given in Section 3. Section 4 reports comparisons with the existing control charts using out-of-control average run length for various shifts as well as simulated real-life data examples. Finally, some conclusions are drawn in Section 5.

2. The proposed nonparametric CEWMA sign chart

Suppose that in a production process, a certain quality characteristic $Y$ has a median $\eta$ as a target value. Moreover, assume that $Z = Y - \eta$ denotes the process deviation from the target value with probability $p = P(Z > 0)$. The proposed chart is based on the sign test in which we test the hypothesis that the probability of a plus sign of $Z$ equals the probability of a minus sign, which is equivalent to the test hypothesis that the process has a specific target value, say $\eta$. Therefore, the process is said to be IC if $p = p_0 = 0.5$; otherwise, the process is deemed to be Out-Of-Control (OOC), i.e., $p = p_1 \neq p_0$. Let \( \{Y_{it}, 1 \leq i \leq n\} \) be an independently and identically distributed sample taken at time $t$ from $Y$ to investigate the variation from the process target value $\eta$. Then, we define:

$$Z_{it} = Y_{it} - \eta \quad \text{and} \quad I_{it} = \begin{cases} 1, & \text{if } Z_{it} > 0 \\ 0, & \text{otherwise} \end{cases}$$

for $i = 1, 2, \ldots$ and $t = 1, 2, \ldots$
Let $S_t$ be the number of observations for which $Z_i$, $i = 1, 2, \ldots, n$, is greater than zero; then, the statistic $S_t = \sum_{i=1}^{n} I_{i,t}$ follows the binomial distribution with parameters $(n, 0.50)$ for an IC process. Based on this assumption, the CEWMA sign chart is proposed by combining the two EWMA statistics as follows:

$$HE_{S_t} = \lambda_1 E_{S_t} + (1 - \lambda_1) HE_{S_{t-1}}, \quad 0 < \lambda_1 \leq 1,$$

$$E_{S_t} = \lambda_2 S_t + (1 - \lambda_2) E_{S_{t-1}}, \quad 0 < \lambda_2 \leq 1,$$

where $S_t$ is the $t$th successively recorded number of $Z_i$ ($> 0$). Note that the mean value of $S_t$ is taken as the initial value of the EWMA statistics, i.e., $HE_{S_0} = E_{S_0} = \mu = n p_0 = n / 2$ and $HE_{S_t}$ is the monitoring statistic of the proposed CEWMA control chart. By expanding Eqs. (1) and (2), we obtain:

$$HE_{S_t} = (1 - \lambda_1)^t HE_{S_0} + \lambda_1 \sum_{i=0}^{t-1} (1 - \lambda_1)^i E_{S_{t-i}},$$

and then by substituting Eq. (4) into Eq. (3), we get:

$$HE_{S_t} = (1 - \lambda_1)^t \mu + \lambda_1 \sum_{i=0}^{t-1} (1 - \lambda_1)^i (1 - \lambda_2)^{t-i} \mu + \lambda_2 \sum_{j=0}^{t-1} (1 - \lambda_2)^i S_{t-j}, \quad (5)$$

$$= (1 - \lambda_1)^t (n p_0) + \lambda_1 \sum_{i=0}^{t-1} (1 - \lambda_1)^i (1 - \lambda_2)^{t-i} (n p_0) + \lambda_2 \sum_{j=0}^{t-1} (1 - \lambda_2)^i S_{t-j}, \quad (6)$$

After some algebraic computation of Eq. (6), we can obtain the mean of $HE_{S_t}$ as:

$$E(HE_{S_t}) = n p_0.$$

Now, by definition, the variance of $HE_{S_t}$ can be obtained as follows:

$$V(HE_{S_t}) = (\lambda_1 \lambda_2)^2 \sum_{i=1}^{t} (1 - \lambda_1)^{2(i-1)}$$

$$\left(\sum_{j=0}^{i-1} \frac{(1 - \lambda_1)^j}{1 - \lambda_2}\right)^2 V(S_t),$$

where $V(S_t) = n p_0 (1 - p_0)$. After some algebraic computation, we obtain the variance of $HE_{S_t}$ as:

$$V(HE_{S_t}) = \left(\frac{\lambda_1 \lambda_2}{\lambda_1 - \lambda_2}\right)^2 \left[\sum_{i=1}^{t} \left(1 - \lambda_1\right)^{2i} \left(1 - \left(1 - \lambda_1\right)^2\right) \frac{1}{1 - \left(1 - \lambda_1\right)^2} \right]$$

$$- \frac{2(1 - \lambda_1) (1 - \lambda_2) \{1 - (1 - \lambda_1)^2 (1 - \lambda_2)^2\}}{1 - (1 - \lambda_1) (1 - \lambda_2)} \{n p_0 (1 - p_0)\}.$$

For $\lambda_1 = \lambda_2 = \lambda$, Eq. (5) reduces to:

$$HE_{S_t} = (1 + \lambda t) (1 - \lambda)^t (n p_0) + \lambda^2 \sum_{i=1}^{t} (t - i + 1) (1 - \lambda)^{t-i} S_i.$$

By definition of variance, we obtain:

$$V(HE_{S_t}) = \lambda^4 \sum_{i=1}^{t} (t - i + 1)^2 (1 - \lambda)^{2(t-i)} V(S_t),$$

which yields:

$$V(HE_{S_t}) = \lambda^4 \sum_{i=1}^{t} (1 - \lambda)^{2(i-1)} \{n p_0 (1 - p_0)\},$$

and after some algebraic work, we finally obtain Eq. (10) as shown in Box I for the variance of Eq. (9). Now, the center line and control limits of the CEWMA sign chart are computed as follows:

$$UCL = n p_0 + k \sqrt{V(HE_{S_t})}, \quad (11)$$

$$CL = n p_0, \quad (12)$$

$$LCL = n p_0 - k \sqrt{V(HE_{S_t})}, \quad (13)$$

where $k$ denotes the distance of the control limits from the center line in multiples of the standard deviation of the CEWMA statistic ($HE_{S_t}$). The three parameters $\lambda_1$, $\lambda_2$, and $k$ are chosen in such a way that a specific IC Average Run Length (ARL) is achieved. Note that if any $HE_{S_t} \geq UCL$ or $HE_{S_t} \leq LCL$, the process is considered to be OOC. Furthermore, if we select $\lambda_1 = 1$ or $\lambda_2 = 1$, the CEWMA sign control chart reduces to the EWMA sign chart proposed by Yang et al. [13].
3. Performance evaluation

The ARL is commonly used to assess the performance of a control chart. Recall that the ARL is the average number of points plotted on a control chart before an OOC signal is shown. In this study, the Monte Carlo simulation is employed to calculate the coefficient values $k$ for various combinations of $\lambda_1$ and $\lambda_2$ after fixing $ARL_0 = 370$ and $n = 10$. To get the average results, we use 100,000 repetitions for each Monte Carlo simulation. Since the distribution of $S_i$ in CEWMA sign statistic follows a binomial distribution, the exact $ARL_0 = 370$ is not always attained. We, therefore, find the $k$ values such that the $ARL_k$ is approximately equal to 370 (i.e., $ARL_k \approx 370$). Some choices of design parameters of the CEWMA sign chart under a nominal $ARL_0$ value of 370 are given in Table 1.

The control limits of the proposed chart depend on $n$, $k$, $\lambda_1$, and $\lambda_2$. It can be observed from Table 1 that the values of $k$ are symmetrical with respect to $\lambda_1$ and $\lambda_2$. There is an increasing trend in $k$ by increasing either $\lambda_1$, $\lambda_2$ or both, except for larger values of $\lambda_1$ and $\lambda_2$. The same strategy can be applied to find the values of $k$ for various $n$’s. In addition to the $ARL_0$, the OOC run length ($ARL_1$) is also calculated which is expected to be adequately small (for a fixed $ARL_0$) to detect the process shift quickly. Considering the IC process proportion $p_0 = 0.50$ and $\lambda_1 = \lambda_2 = \lambda = 0.05$, the $ARL_1$ values of the CEWMA sign chart are computed in Table 2 by using different values of $n$ and the expected OOC process proportion ($p_1$). Besides, the performance of the CEWMA sign chart is also investigated for $\lambda_1 \neq \lambda_2$. In this scenario, the ARL values for smoothing constants ($\lambda_1 = 0.05$, $\lambda_2 = 0.10$) and ($\lambda_1 = 0.05$, $\lambda_2 = 0.20$) under various $n$ and $p_1$ are presented in Tables 3 and 4, respectively. When $p = 0.50$, the attained $ARL_0$ values are approximately equal to the desired value of 370. From Tables 2-4, it can be noticed that $ARL_1$ values are inversely related to $p_1 = 0.50$ and $n$. It can also be observed that $k$ changes marginally as the sample size $n$ increases. Moreover, for smaller shifts, the $ARL_1$ values increase marginally with increase in one of the smoothing constants $\lambda_1$ and $\lambda_2$.

In the following, we further investigate the performance of the proposed CEWMA control chart under the assumption that the critical quality characteristic of a process follows some specific distribution. The performance of the proposed CEWMA control chart under various symmetrical and asymmetrical distributions is assessed by using different run-length characteristics such as ARL, median run length ($MDRL$), and Standard Deviation Run Length ($SDRL$). To be more precise, by setting $ARL_0 = 370$, the run length distribution for $\lambda_1 = \lambda_2 = 0.05$ and $n = 10$ is computed under the standard normal distribution, $N(0,1)$; the Student’s $t$ distribution with 4 and 8 degrees of freedom; the heavy-tailed symmetric distributions that include the Laplace (or double exponential) distribution $Laplace(0,1/\sqrt{2})$, the Logistic distribution $LG(0,\sqrt{3}/\pi)$, the Weibull distribution $Weibull(2,1)$, the Gamma distribution Gamma$(4,1)$, and the contaminated normal (CN) distribution. Recall that the CN distribution is obtained by combining two normal distributions $N(0,\sigma_1^2)$ and $N(0,\sigma_2^2)$ using the relation $(1-\varphi)N(0,\sigma_1^2) + \varphi N(0,\sigma_2^2)$, where $\varphi$ denotes the proportion of contamination. This distribution is generally used to evaluate the outlier resistance power of the control charts. In our study, we set $\sigma_2^2 = 4\sigma_1^2$ and the proportion of contamination $\varphi = 0.10$. To

| $\lambda_1$ | 0.05 | 0.10 | 0.20 | 0.25 | 0.30 | 0.40 | 0.50 | 0.60 | 0.70 | 0.80 | 0.90 |
|-------------|------|------|------|------|------|------|------|------|------|------|------|
| $\lambda_2$ |      |      |      |      |      |      |      |      |      |      |      |
| 0.05        | 1.954| 2.092| 2.227| 2.271| 2.304| 2.353| 2.392| 2.428| 2.458| 2.505|
| 0.10        | 2.092| 2.245| 2.387| 2.429| 2.468| 2.526| 2.562| 2.598| 2.629| 2.66|
| 0.20        | 2.227| 2.387| 2.528| 2.571| 2.61 | 2.666| 2.711| 2.744| 2.772| 2.795|
| 0.25        | 2.271| 2.429| 2.571| 2.619| 2.655| 2.71 | 2.751| 2.784| 2.809| 2.832|
| 0.30        | 2.304| 2.468| 2.61 | 2.655| 2.692| 2.743| 2.782| 2.811| 2.837| 2.855|
| 0.40        | 2.353| 2.526| 2.666| 2.71 | 2.743| 2.792| 2.825| 2.853| 2.871| 2.883|
| 0.50        | 2.392| 2.562| 2.711| 2.751| 2.782| 2.825| 2.856| 2.875| 2.884| 2.889|
| 0.60        | 2.428| 2.598| 2.744| 2.784| 2.811| 2.853| 2.875| 2.884| 2.893| 2.885|
| 0.70        | 2.455| 2.629| 2.772| 2.809| 2.837| 2.871| 2.884| 2.89 | 2.894| 2.882|
| 0.80        | 2.481| 2.66 | 2.795| 2.832| 2.855| 2.878| 2.889| 2.893| 2.882| 2.868|
| 0.90        | 2.505| 2.682| 2.822| 2.852| 2.871| 2.883| 2.892| 2.885| 2.868| 2.855| 2.82
Table 2. The ARL values of the Composite Exponentially Weighted Moving Average (CEWMA) sign control chart for $\lambda_1 = \lambda_2 = \lambda = 0.05$.

| n  | 9   | 10  | 11  | 12  | 13  | 14  | 15  | 16  | 17  | 18  | 19  | 20  | 21  | 22  | 23  | 24  | 25  |
|----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|
| k  | 1.956 | 1.954 | 1.962 | 1.955 | 1.953 | 1.961 | 1.958 | 1.966 | 1.956 | 1.955 | 1.963 | 1.957 | 1.964 | 1.957 | 1.955 | 1.964 | 1.957 |
| 0.05 | 1.1 | 1.1 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 |
| 0.10 | 1.3 | 1.3 | 1.1 | 1.1 | 1.1 | 1.1 | 1.1 | 1.1 | 1.1 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 |
| 0.15 | 1.6 | 1.6 | 1.3 | 1.3 | 1.4 | 1.2 | 1.2 | 1.1 | 1.1 | 1.1 | 1.1 | 1.1 | 1.1 | 1.0 | 1.0 | 1.0 | 1.0 |
| 0.20 | 2.0 | 2.0 | 1.6 | 1.6 | 1.7 | 1.4 | 1.4 | 1.3 | 1.3 | 1.2 | 1.2 | 1.3 | 1.2 | 1.1 | 1.2 | 1.1 |
| 0.25 | 2.8 | 2.6 | 2.3 | 2.2 | 2.2 | 1.9 | 1.9 | 1.7 | 1.7 | 1.6 | 1.7 | 1.5 | 1.6 | 1.4 | 1.4 | 1.3 | 1.3 |
| 0.30 | 4.1 | 3.8 | 3.4 | 3.4 | 3.2 | 2.8 | 2.7 | 2.4 | 2.3 | 2.2 | 2.1 | 2.1 | 2.1 | 2.0 | 1.9 | 1.8 | 1.8 |
| 0.35 | 7.0 | 6.3 | 5.7 | 5.4 | 5.1 | 4.7 | 4.4 | 4.1 | 3.9 | 3.6 | 3.5 | 3.5 | 3.3 | 3.1 | 2.9 | 2.9 |
| 0.40 | 13.8 | 12.8 | 11.5 | 10.9 | 10.3 | 9.5 | 8.9 | 8.3 | 8.0 | 7.8 | 7.2 | 7.0 | 6.9 | 6.3 | 6.3 | 6.0 | 5.9 |
| 0.45 | 42.1 | 38.6 | 35.4 | 33.6 | 31.6 | 29.7 | 28.0 | 26.2 | 25.3 | 24.7 | 23.0 | 22.3 | 21.7 | 20.5 | 20.0 | 18.9 | 18.6 |
| p  | 50.0 | 371.0 | 370.8 | 370.5 | 369.5 | 371.8 | 371.7 | 371.5 | 370.4 | 371.4 | 371.3 | 371.9 | 370.1 | 372.8 | 369.4 | 370.9 | 371.6 |

Table 3. The ARL values of the Composite Exponentially Weighted Moving Average (CEWMA) sign control chart for $\lambda_1 = 0.05$ and $\lambda_2 = 0.10$.

| n  | 9   | 10  | 11  | 12  | 13  | 14  | 15  | 16  | 17  | 18  | 19  | 20  | 21  | 22  | 23  | 24  | 25  |
|----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|-----|
| k  | 2.097 | 2.092 | 2.105 | 2.098 | 2.10 | 2.101 | 2.098 | 2.096 | 2.101 | 2.098 | 2.096 | 2.098 | 2.095 | 2.101 | 2.097 | 2.095 | 2.096 |
| 0.05 | 1.1 | 1.1 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 |
| 0.10 | 1.3 | 1.3 | 1.1 | 1.1 | 1.1 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 |
| 0.15 | 1.6 | 1.5 | 1.3 | 1.3 | 1.3 | 1.2 | 1.2 | 1.2 | 1.1 | 1.1 | 1.1 | 1.1 | 1.0 | 1.0 | 1.0 | 1.0 | 1.0 |
| 0.20 | 2.0 | 2.1 | 1.7 | 1.7 | 1.7 | 1.4 | 1.4 | 1.3 | 1.3 | 1.3 | 1.2 | 1.2 | 1.2 | 1.1 | 1.2 | 1.1 |
| 0.25 | 3.0 | 2.8 | 2.4 | 2.4 | 2.2 | 2.0 | 2.0 | 2.0 | 2.0 | 1.7 | 1.7 | 1.7 | 1.6 | 1.6 | 1.4 | 1.4 | 1.4 |
| 0.30 | 4.1 | 4.1 | 3.6 | 3.4 | 3.3 | 2.9 | 2.9 | 2.8 | 2.5 | 2.4 | 2.3 | 2.4 | 2.3 | 2.2 | 2.0 | 2.0 | 1.9 |
| 0.35 | 7.4 | 6.9 | 6.1 | 5.9 | 5.4 | 4.9 | 4.7 | 4.6 | 4.2 | 4.1 | 4.0 | 3.7 | 3.6 | 3.3 | 3.3 | 3.3 | 3.1 |
| 0.40 | 14.7 | 13.4 | 12.3 | 11.6 | 10.8 | 10.0 | 9.6 | 9.2 | 8.6 | 8.2 | 7.9 | 7.5 | 7.2 | 6.8 | 6.6 | 6.5 | 6.2 |
| 0.45 | 44.9 | 41.9 | 36.7 | 35.7 | 33.9 | 31.0 | 29.7 | 28.8 | 26.8 | 25.9 | 24.7 | 23.4 | 22.8 | 21.5 | 21.1 | 20.7 | 19.7 |
| p  | 50.0 | 372.5 | 370.9 | 371.0 | 370.9 | 370.9 | 370.1 | 371.5 | 369.7 | 370.7 | 371.2 | 371.3 | 370.3 | 369.9 | 371.5 | 371.9 | 370.7 | 370.4 |

The ARLq values remain the same under all continuous symmetrical and asymmetrical distributions which validate the theory of nonparametric control charts. Furthermore, the ARL values decline rapidly with increase in the size of the shift for all the distributions determined by the OOC run length characteristics, the shift in the process mean is introduced in terms of process standard deviation ($\sigma$), i.e., $\mu_1 = \mu_0 + \delta \sigma$ where $\delta$ determines the magnitude of the shift. The results are reported in Table 5. It can be noticed from Table 5 that
Table 4. The ARL values of the Composite Exponentially Weighted Moving Average (CEWMA) sign control chart for \( \lambda_1 = 0.05 \) and \( \lambda_2 = 0.20 \).

| \( p \) | 0.50 | 0.60 | 0.65 | 0.70 | 0.75 | 0.80 | 0.85 | 0.90 | 0.95 |
|---|---|---|---|---|---|---|---|---|---|
| \( ARL_0 \) | 371.2 | 369.4 | 370.8 | 370.3 | 370.7 | 370.3 | 370.5 | 369.2 | 370.7 |
| \( ARL_1 \) | 371.4 | 370.8 | 370.3 | 370.7 | 370.3 | 370.5 | 369.2 | 370.7 | 371.1 |

4. Performance comparison

In this section, the performance of the proposed chart is compared with those of the existing nonparametric EWMA sign and CUMSUM mean control chart by using \( ARL_1 \) values for various shifts. Besides, two examples are presented in this section. Considering \( n = 10, 15, 20, p_0 = 0.50, \) and \( ARL_0 = 370, \) the \( ARL_1 \) values of the nonparametric EWMA sign, CUMSUM mean, and CEWMA sign control charts are shown in Figures 1–3 for various shifts in the process proportion.

![Figure 1](image1.png)

**Figure 1.** ARL curves of the nonparametric CEWMA sign, EWMA sign, and CUMSUM mean control charts at \( n = 10 \) and \( ARL_0 \approx 370. \)

![Figure 2](image2.png)

**Figure 2.** ARL curves of the nonparametric CEWMA sign, EWMA sign, and CUMSUM mean control charts at \( n = 15 \) and \( ARL_0 \approx 370. \)

![Figure 3](image3.png)

**Figure 3.** ARL curves of the nonparametric CEWMA sign, EWMA sign, and CUMSUM mean control charts at \( n = 20 \) and \( ARL_0 \approx 370. \)
Table 5. The run-length characteristic of the proposed Composite Exponentially Weighted Moving Average (CEWMA) chart under various distributions for $\lambda_1 = \lambda_2 = 0.05$, $n = 10$, and $k = 1.954$ at $ARL_0 \approx 370$.

| Distribution | Characteristic | 0.00 | 0.05 | 0.10 | 0.25 | 0.50 | 0.75 | 1.00 | 1.50 | 2.00 | 2.5 | 3.0 |
|-------------|---------------|------|------|------|------|------|------|------|------|------|------|------|
| $N(0, 1)$   | ARL           | 370.8| 141.9| 54.2 | 12.9 | 4.1  | 2.2  | 1.6  | 1.1  | 1.0  | 1.0  | 1.0  |
|             | SDRL          | 423.8| 150.8| 50.6 | 10.4 | 3.0  | 1.3  | 0.7  | 0.4  | 0.1  | 0.0  | 0.0  |
|             | MDRL          | 235  | 97   | 42   | 10   | 3    | 2    | 1    | 1    | 1    | 1    | 1    |
| $t(4)$      | ARL           | 371.4| 99.3 | 35.4 | 8.2  | 2.8  | 1.7  | 1.4  | 1.1  | 1.0  | 1.0  | 1.0  |
|             | SDRL          | 416.3| 100.7| 31.4 | 6.4  | 1.8  | 0.8  | 0.5  | 0.3  | 0.2  | 0.1  | 0.0  |
|             | MDRL          | 239  | 71   | 28   | 7    | 2    | 1    | 1    | 1    | 1    | 1    | 1    |
| $t(8)$      | ARL           | 371.1| 126.4| 45.7 | 10.9 | 3.5  | 2.0  | 1.5  | 1.1  | 1.0  | 1.0  | 1.0  |
|             | SDRL          | 417.9| 129.7| 41.9 | 8.7  | 2.5  | 1.1  | 0.7  | 0.3  | 0.2  | 0.1  | 0.0  |
|             | MDRL          | 237  | 88   | 36   | 9    | 3    | 2    | 1    | 1    | 1    | 1    | 1    |
| $CN$        | ARL           | 370.3| 128.1| 47.2 | 11.1 | 3.6  | 2.1  | 1.5  | 1.1  | 1.0  | 1.0  | 1.0  |
|             | SDRL          | 417.3| 133.4| 42.9 | 8.9  | 2.6  | 1.2  | 0.7  | 0.3  | 0.2  | 0.1  | 0.0  |
|             | MDRL          | 234  | 88   | 37   | 9    | 3    | 2    | 1    | 1    | 1    | 1    | 1    |
| $LG(0, \sqrt{3}/\pi)$ | ARL    | 368.4| 123.2| 45.0 | 10.6 | 3.4  | 2.0  | 1.5  | 1.1  | 1.0  | 1.0  | 1.0  |
|             | SDRL          | 417.8| 127.7| 40.7 | 8.5  | 2.4  | 1.1  | 0.7  | 0.3  | 0.2  | 0.1  | 0.0  |
|             | MDRL          | 235  | 86   | 36   | 9    | 2    | 2    | 1    | 1    | 1    | 1    | 1    |
| $Laplace(0, 1/\sqrt{2})$ | ARL    | 309.2| 69.8 | 25.1 | 6.4  | 2.6  | 1.7  | 1.4  | 1.1  | 1.0  | 1.0  | 1.0  |
|             | SDRL          | 412.8| 67.2 | 21.2 | 4.9  | 1.6  | 0.8  | 0.6  | 0.3  | 0.2  | 0.1  | 0.1  |
|             | MDRL          | 237  | 52   | 20   | 5    | 2    | 2    | 1    | 1    | 1    | 1    | 1    |
| Gamma(4, 1) | ARL           | 371.6| 132.7| 50.5 | 12.4 | 4.1  | 2.3  | 1.7  | 1.2  | 1.0  | 1.0  | 1.0  |
|             | SDRL          | 415.9| 140.3| 46.4 | 10.0 | 3.0  | 1.4  | 0.8  | 0.4  | 0.2  | 0.1  | 0.0  |
|             | MDRL          | 237  | 91   | 39   | 10   | 3    | 2    | 2    | 1    | 1    | 1    | 1    |
| Weibull(2, 1) | ARL    | 309.7| 165.1| 68.5 | 17.5 | 6.5  | 3.9  | 2.8  | 1.9  | 1.6  | 1.4  | 1.3  |
|             | SDRL          | 415.6| 174.6| 64.9 | 14.3 | 5.1  | 2.8  | 1.8  | 1.0  | 0.7  | 0.6  | 0.5  |
|             | MDRL          | 237  | 113  | 52   | 14   | 5    | 3    | 2    | 2    | 1    | 1    | 1    |

It can be observed that the nonparametric EWMA sign and CUSUM mean control charts produce almost similar results and the proposed CEWMA sign control chart remains on the lower side for all shifts. This shows that the proposed chart outperforms for all shifts of $p_1$ and for different sample sizes.

### 4.1. Example 1

To illustrate the shift detection ability of the CEWMA sign and EWMA sign control charts, the following example with description taken from Montgomery [36] is considered.

"The fill volume of soft-drink beverage bottles is an important quality characteristic. The volume is measured (approximately) by placing a gauge over the crown and comparing the height of the liquid in the neck of the bottle against a coded scale. On this scale, a reading of zero corresponds to the correct fill height."

In this example, we have 15 subgroups each with size of 10 and the target value of 0. By choosing $\lambda_1 = 0.05$ and $k = 2.49$ at $ARL_0 \approx 370$, we obtain 4.36 and 5.63 as the lower and upper control limits for the EWMA sign chart, respectively. The monitoring statistic $E_{S_i}$ of the EWMA sign chart is computed and plotted against their respective control limits in Figure 4. We see that the nonparametric EWMA sign chart detects an OOC signal in sample 13. For rational evaluation of the EWMA sign scheme, the proposed nonparametric CEWMA sign chart is constructed by considering $\lambda_1 = \lambda_2 = \lambda = 0.05$ and $k = 1.954$, as shown in Figure 5. The last two columns of Table 6 show time-varying lower and upper control limits of
Table 6. Existing nonparametric Exponentially Weighted Moving Average (EWMA) sign and the proposed nonparametric Composite Exponentially Weighted Moving Average (CEWMA) sign control charts for the dataset from Montgomery [36], the fill volume of soft-drink beverage bottle.

| Sample | $X_1$ | $X_2$ | $X_3$ | $X_4$ | $X_5$ | $X_6$ | $X_7$ | $X_8$ | $X_9$ | $X_{10}$ | $S_i$ | $E_{Si}$ | $HE_{Si}$ | LCL | UCL |
|--------|------|------|------|------|------|------|------|------|------|-------|------|---------|---------|-----|-----|
| 1      | 2.5  | 0.5  | 2    | -1   | 1    | -1   | 0.5  | 1.5  | 0.5  | -1.5  | 5    | 10.00   | 50.00   | 49.923| 50.077|
| 2      | 0    | 0    | 0.5  | 1    | 1.5  | 1    | -1   | 1    | 1.5  | -1.5  | 5    | 14.50   | 51.00   | 49.834| 50.166|
| 3      | 1.5  | 1    | 1    | -1   | 1    | 0    | -1.5 | 0    | -1   | 1     | 4    | 5.0878  | 50.18   | 49.733| 50.267|
| 4      | 0    | 0.5  | -2   | 0    | -1   | 1.5  | -1.5 | 0    | -2   | -1.5  | 2    | 4.9334  | 50.117  | 49.964| 50.076|
| 5      | 0    | 0    | 0    | -0.5 | 0.5  | 0    | -0.5 | 0    | 0    | 0     | 2    | 4.7867  | 50.004  | 49.9510| 50.0490|
| 6      | 1    | -0.5 | 0    | 0    | 0.5  | -1   | 1    | 1    | -2   | 1     | 4    | 4.7474  | 49.978  | 49.9393| 50.0607|
| 7      | 1    | -1   | -1   | -1   | 1.5  | 0    | 1    | 0    | 0    | 0     | 3    | 4.6600  | 49.914  | 49.9274| 50.0726|
| 8      | 0    | -1.5 | -0.5 | 1.5  | 0    | 0    | 0    | -1   | 0.5  | -0.5  | 2    | 4.5270  | 49.942  | 49.9156| 50.0844|
| 9      | -2   | -1.5 | 1.5  | 1    | 0    | 0    | 0.5  | 1    | 0    | 1     | 5    | 4.5506  | 49.929  | 49.9038| 50.0662|
| 10     | -0.5 | 3.5  | 0    | -1   | -1.5 | -1.5 | 0    | -1   | 1    | 0.5   | 3    | 4.4731  | 49.9064 | 49.8922| 50.1078|
| 11     | 0    | 1.5  | 0    | 0    | 2    | -1.5 | 0.5  | -0.5 | 2    | -1    | 4    | 4.4195  | 48.8836 | 48.8608| 51.1192|
| 12     | 0    | -2   | -0.5 | 0    | -0.5 | 2    | 1.5  | 0    | 0.5  | -1    | 3    | 4.3770  | 48.8582 | 48.8696| 51.3004|
| 13     | -1   | -0.5 | -0.5 | -1   | 0    | 0.5  | 0.5  | -1.5 | -1   | 1     | 2    | 4.2581  | 48.8282 | 48.8588| 51.1412|
| 14     | 0.5  | 1    | -1   | -0.5 | -2   | -1   | -1.5 | 0    | 1.5  | 1.5   | 4    | 4.2452  | 47.9991 | 48.8483| 51.5157|
| 15     | 1    | 0    | 1.5  | 1.5  | 1    | -1   | 0    | 1    | -2   | -1.5  | 5    | 4.2830  | 47.7333 | 48.8381| 51.6119|
charts are provided in Table 7 and displayed graphically in Figures 8–10. In these schemes, the statistic $E_S$ is plotted against the control limits $LCL = 6.7279$ and $UCL = 8.2721$ for the nonparametric EWMA sign chart, $C^+$'s and $C^-$'s are plotted against the decision interval $H_1$ for the nonparametric CUSUM mean chart, and $HE_S$ is plotted against time-varying control limits for the proposed CEWMA sign chart.

It is apparent from Figure 10 that the suggested control chart triggers OOC signals in samples 35–40, whereas Figures 6–9 show that the Shewhart $\bar{X}$, classical EWMA, nonparametric EWMA sign, and nonparametric CUSUM mean control charts fail to identify this shift in the given dataset. This example points to the superiority of the proposed chart in identifying small shifts in the process over the four other competing parametric and nonparametric control charts.

5. Conclusion

In this study, a new nonparametric Composite Exponentially Weighted Moving Average (CEWMA) sign chart that represented the combination of two EWMA statistics was proposed. The proposed chart could be used to detect process shifts quickly, particularly smaller shifts, by using different combinations of the sensitivity parameters $\lambda_1$ and $\lambda_2$. The suggested CEWMA sign chart could attract wider applicability as the performance of the CEWMA sign chart was superior to the existing nonparametric Exponentially Weighted Moving Average (EWMA) sign chart and the nonparametric CUSUM mean chart. Furthermore, the nonparametric EWMA sign chart is a special case of the proposed CEWMA chart when either $\lambda_1$ or $\lambda_2$ equals one. The practical implementation of our proposed chart was illustrated using a real-life example. Moreover, results based on the simulation study showed that the suggested chart completely dominated the
other competing parametric and nonparametric control charts in detecting smaller shifts. Hence, the use of CEWMA sign control chart is recommended for the practitioners, if complete knowledge about the underlying process distribution is not available. One can further extend this scheme by developing the CEWMA sign chart using other nonparametric tests such as the Mann-Whitney test.

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