Cohort Study

Indwelling catheters increase altered mental status and urinary tract infection risk: A retrospective Cohort Study

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ARTICLE INFO

Keywords:
Urinary catheter
Perioperative complication
Delirium
Urinary tract infection

ABSTRACT

Background: Although indwelling urinary catheters (IUCs) are used intraoperatively and may cause complications (e.g., delirium), only few robust studies have investigated the association between intraoperative IUC use and complications. We hypothesized that IUC use might increase the postoperative incidence of altered mental status and/or urinary catheter infection.

Materials and methods: In this retrospective single-center cohort study, we analyzed the data of adult patients undergoing surgery at our facility between January 2013 and December 2018. The primary endpoint was altered mental status and/or incidence of urinary catheter infections. The patients were divided into IUC and control groups. A multivariable logistic regression model was used to identify the predictors of postoperative complications, and a multivariable Cox proportional hazards regression model was used to analyze hospital discharge in unmatched and inverse propensity-weighted patients.

Results: Of the 14,284 patients that were reviewed, we analyzed 5112 patients (control group, 44.0%; IUC group, 56.0%). Almost all procedures comprised less invasive surgeries. The prevalence of postoperative altered mental status and postoperative urinary catheter infection were 3.56% and 0.04%, respectively. After inverse propensity weighting, all baseline characteristics were similar between the two groups. However, patients with IUCs had a higher risk of postoperative complications (adjusted odds ratio, 1.97; 95% confidence interval [CI], 1.50–2.59) and prolonged hospital stays (hazard ratio, 0.84; 95% CI, 0.80–0.89).

Conclusion: In patients undergoing less invasive surgery, IUCs may be associated with a relatively high risk of altered mental status or urinary catheter infection. These data may facilitate preoperative discussions regarding the perioperative use of IUCs.

1. Introduction

During surgery, indwelling urinary catheters (IUCs) are frequently used to assess urine output as an indicator for the optimal blood pressure (1–7) and a predictor of postoperative acute kidney injury (AKI) (8–10). Worldwide, nearly 80% of all anesthesiologists measure the urine output (11,12–17). However, IUCs can cause complications, most frequently catheter-induced urinary tract infections. Approximately 12–16% of adult hospital inpatients will require an IUC, and a patient has a 3–7% increased risk of acquiring a catheter-associated urinary tract infection each day the IUC remains (18). The Center for Disease Control and Prevention (CDC) considers an IUC indication appropriate when one or more of the following criteria are met: a patient undergoing urologic or other surgery on contiguous structures of the genitourinary tract;
anticipated prolonged surgery duration, catheters inserted for this reason should be removed in the post-anesthesia care unit; anticipated need for large-volume infusions or diuretics during surgery; or need for intraoperative measuring of the urinary output [18]. However, these suggestions were not based on robust evidence evaluating risks (e.g., delirium, urinary tract infection) and benefits (e.g., maintaining optimal blood pressure, decreased incidence of postoperative AKI) of intraoperative IUC use.

Some studies reported that perioperative IUC usage may induce discomfort, delirium, or urinary tract infection [19–24]. Almost all studies examining perioperative IUC usage had low evidence levels because delirium and urinary tract infection were not primary endpoints. Eide et al. [22] demonstrated that urinary catheter use is associated with delirium. However, their study population comprised only patients undergoing transcatheter aortic valve implantation [22]. Therefore, the benefit-risk assessment of intraoperative IUC usage is inconclusive.

Kuriyama et al. [25] reported that only 5.3% of the elective surgical patients from seven participating intensive care units (ICUs) in Japan had an appropriate IUC based on indications suggested by the CDC. Eide et al. [22] also suggested that studies should evaluate whether IUC use could be avoided entirely in patients with transcatheter aortic valve implantation. Therefore, we hypothesized that perioperative IUC use might increase the incidence of altered mental status and/or urinary tract infection postoperatively but might not decrease that of postoperative AKI. Before we conduct a randomized controlled trial (RCT) to test this hypothesis, we supposed that we should estimate the incidence of altered mental status or urinary tract infection postoperatively associated with IUCs. For this reason, we conducted the current retrospective cohort study that may be helpful to determine whether intraoperative IUC use is predominantly a risk or a benefit.

2. Methods

The Institutional Review Board of our institution approved the study and waived the requirement for informed consent because the procedures comprised no additional interventions and because all patient data were anonymous. This retrospective cohort study was registered with UMIN-ICDR (UMIN000042664, https://upload.umin.ac.jp/cgi-bin/ctr/ctr_view_reg.cgi?recptno=R000048679) in accordance with the declaration of Helsinki. The opt-out method was available on the hospital’s website. This study was conducted following the Strengthening the Reporting of Cohort Studies in Surgery (STROCSS) [26] and the Strengthening the Reporting of Observational Studies in Epidemiology (STROBE) guidelines. We included all adult (age ≥18 y) noncardiac surgical patients who required general anesthesia between January 1, 2013, and December 31, 2018. We excluded the following patients: patients undergoing urology surgery because IUCs were needed during and after surgery (to assess complications, urine output, and to prevent complications), patients who needed an IUC during surgery (to assess the urine output, the indicators for optimal blood pressure were as follows: emergency surgery, mechanical ventilation before surgery, and carotid endarterectomy; for assessing urinary drainage because of urination disorder, the indicators were as follows: traumatic spinal cord injury, epidural catheter use, spinal anesthesia, severe motor and intellectual disabilities), patients undergoing neurosurgery affecting postoperative consciousness and immobility (cereotomy, trepanation, intracranial endovascular therapy), patients who could not speak and who could not vocalize unfamiliar words (tracheostomy), patients undergoing chronic hemodialysis due to reduced urine output, and patients who experienced cardiac arrest before surgery.

Patient characteristics and intraoperative data were retrieved from the operating room database (ORSYS, Philips Electronics Japan, Tokyo, Japan) and the patient information system (PIMS, Philips Electronics Japan, Tokyo, Japan; HOPE EGMAIN-GX, FUJITSU, Kanagawa, Japan) of the hospital. We did not collect post-hospital discharge data because those were not recorded routinely.

The primary endpoint was the incidence of altered mental status and/or urinary tract infection. The secondary endpoints were the length of postoperative hospital stay and AKI incidence after surgery.

We use the expression “altered mental status” instead of delirium because we usually do not employ delirium diagnostic tools such as the Confusion Assessment Method for the ICU (CAM-ICU); thus, this information was not available. Before and after conducting this study, the descriptions of altered mental status in Japanese were discussed by seven anesthesiologists to decide whether these descriptions reflected altered mental status and not an arbitrary selection. Nursing records between the time of discharge from the operating room and the next day at 23:59 were analyzed.

The definition of a urinary tract infection was based on the CDC guidelines [27]. A patient had to meet both of the following criteria: presence of at least one of the following signs or symptoms: fever (> 38.0°C), suprapubic tenderness, costovertebral angle pain or tenderness, urinary urgency, urinary frequency, and dysuria and identification of no more than two species of organisms from urine culture, at least one of which was a bacterium with a load of ≥10^6 colony-forming units/mL. Postoperative AKI was defined according to the Kidney Disease Improving Global Outcomes guidelines [8]. Severity levels were defined based on changes in serum creatinine levels: stage 1, increase in serum creatinine levels (sCr) of >0.3 mg/dL within 48 h or an increase corresponding to 1.5–1.9-times the baseline value within 7 d; stage 2, increase in sCr corresponding to 2.0–2.9-times the baseline value; and stage 3, increase in sCr of ≥4.0 mg/dL, an increase corresponding to >3.0-times the baseline value, or the initiation of or increase in the dose of renal replacement therapy within 7 d. Urine output criteria were not used in the present study.

2.1. Sample size

Our sample size estimate was based on the expected incidence of postoperative altered mental status or urinary tract infection as suggested in previous studies [18,20,24,28–31]. Assuming a baseline incidence of the primary endpoints of around 20% in the control group and considering a 15% increase in incidence in the IUC group to be clinically important, we estimated that a total of 1892 patients would be required to demonstrate this difference (type 1 error < 0.05 and a power of 0.8). We expected approximately 5000 eligible patients to visit the medical center over a 6-year period.

2.2. Adjustment for differences between groups

All patients were divided into two groups (IUC group and control group) according to intraoperative IUC use. In our institution, the need for IUC during surgery was decided by the physician or anesthesiologist. The urinary catheter was inserted by either a trained physician or by the nurses. We supposed that characteristics would differ between the two groups. Therefore, inverse probability weighting based on the propensity-score was used as the primary tool to adjust for differences [32]. We constructed a logistic regression model to calculate the propensity-score for each subject who underwent intraoperative IUC placement based on risk factors for cardiovascular disease and postoperative AKI [33]. Preoperative variables that would affect the indication for IUC use were as follows: age; sex; body mass index (kg/m²); American Society of Anesthesiology physical status (ASA-PS) score; surgical specialty; the serum creatinine level (mg/dL) before surgery; a history of dual antiplatelet therapy, defined as the use of aspirin or cilostazol combined with a P2Y12 inhibitor; a history of using oral anticoagulant medications, defined as warfarin or direct oral anticoagulants; and mean blood pressure on admission to the operating room (mmHg), a non-invasive blood pressure measurement excluding pressure pulse values < 20 mmHg, which were considered measuring errors. Covariate balances before and after inverse probability weighting were compared.
using standardized differences [33]. Absolute standardized differences (ASD) within 10% for all variables were indicative of successful balancing.

2.3. Missing data imputation

We filled in missing data with multiple imputations using the R package “mice version 3.6.0” (R Core Team, 2019) [27]. To predict missing values, we used information about intervention allocation (with or without IUC use) and the variables independently associated with postoperative hospital stay in the multiple logistic regression analysis.

2.4. Statistical analysis

Baseline characteristics for the full cohort are summarized as numbers and percentages for categorical variables and as median values and interquartile ranges for continuous variables. Postoperative variables before and after inverse probability weighting (IPW) were analyzed for between-group differences using the Wilcoxon-Mann-Whitney test for continuous variables, Fisher’s exact test for categorical variables, and the log-rank test for postoperative hospital stay. A two-tailed P value < 0.05 was considered statistically significant. A multivariable logistic regression model was used to identify the predictors of the primary endpoints, and a multivariable Cox proportional hazards regression model was used for analyzing postoperative hospital stay. We used Akaike’s Information Criteria for goodness-of-fit measures. Kaplan–Meier curves were used to show the incidence of hospital discharge from the day of surgery through 30 d.

Several sensitivity analyses were performed. First, patient matching based on the propensity-score was performed to adjust for differences. Second, we constructed a doubly robust strategy combining IPW with a logistic regression model [34]. All statistical analyses were performed using R (version 3.6.1, R Foundation for Statistical Computing, Vienna, Austria).

3. Results

Between January 1, 2013, and December 31, 2018, 14,284 surgical patients were managed by an anesthesiologist. Of those, 5130 patients met the inclusion criteria. We excluded 18 patients because we could not assess the primary endpoint and we think that the procedures were not associated with IUC use (10 patients received reoperation, and 9 patients in these cases. Patients who had mechanical ventilation would have received an IUC because re-surgery was an emergency and we received mechanical ventilation within 24 h after surgery with 1 patient met the inclusion criteria. We excluded 18 patients because we could not be verified this because we could not determine the day of surgery; thus, IUC-induced discomfort may have caused altered mental status in our study.

Our study verified the hypothesis that IUC use is associated with altered mental status and urinary tract infection.

Urinary tract infection is a common IUC complication. A meta-analysis by Saint [31] indicated a 26% pooled cumulative incidence of developing bacteriuria in patients with IUCs for 2–10 d; this bacteriuria incidence accounted to a 3–10% daily estimate. Our study identified two cases of urinary tract infection exclusively in the IUC group (0.06%). We removed the urinary catheters on postoperative days 6 and 2, respectively. Our IUC durations were similar to those in the study by Saint [30], but the urinary tract infection incidence in our study was considerably lower. We speculate that the mean IUC duration might have been shorter in our study as it included less invasive surgeries. However, we were not able to verify this because we could not determine the day of IUC removal.

In our study, hospital stays were longer in the IUC group than in the control group (Tables 2 and 3; Fig. 3). IUCs may lead to delirium or urinary tract infection. Eide et al. [22] verified this hypothesis in 136 patients, but their study population was limited to patients who underwent aortic valve surgery (incidence ratio of delirium, 55.9%; OR, 1.04; 95% CI, 1.00–1.08). Our study included various types of surgeries, although they were all minimally invasive procedures. Besides, the population of our study was larger than those of previous studies, thereby validating our findings.

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Fig. 1. Patient enrollment procedure and study flow chart.
Abbreviations: GA, general anesthesia; MV, mechanical ventilation; POD, postoperative day.
maintain optimal blood pressure [4,6-8,11-17]. However, there is little evidence on the management of urine output for improving the prognosis (e.g., decreased mortality) [37,38]. In our study, almost all surgeries were less invasive, had short durations (about 100 min), and had low urine outputs (median volume about 150 mL). This study was inconclusive on whether urine output should be measured for subsequent blood pressure regulation. Further studies regarding blood pressure control based on intraoperative urine output are needed.

Intraoperative oliguria for more than 2 h may predict postoperative AKI [9,10], but this is controversial [9,10,25]. In our study, surgery durations were short and AKI incidences were low. Thus, it was difficult to make clinical judgments on intraoperative oliguria and to predict postoperative AKI. We suggest that IUCs may not be needed to predict postoperative AKI in minimally invasive surgical patients.

Baldini et al. [39] recommend bladder catheterization in outpatients when the bladder volume exceeds 600 mL over 120 min [39]. The CDC suggests that IUCs are appropriate for patients with anticipated intraoperative large-volume infusions or diuretics use [18]. In our study, the median urine output was substantially lower than the bladder capacity. Future studies should assess the association between surgery duration and intraoperative IUC use.

This is the largest observational study focusing on IUCs’ association with complications after minimally invasive surgery. However, this study had some limitations. First, the endpoint “altered mental status” was not based on diagnostic criteria or on a delirium tool such as CAM-ICU. Most previous studies assessing the prevalence of delirium used
The use of variable descriptions for “altered mental state” could have affected the reporting of this occurrence. Therefore, the descriptions defining “altered mental status” were discussed before the study started and individually assessed to eliminate bias. Second, most procedures were minimally invasive. Third, the study included a few patients who underwent major abdominal surgery. We mostly use IUCs in patients with epidural catheters because epidural anesthesia may inhibit parasympathetic neurons of the spinal cord causing urinary retention. We excluded abdominal surgery patients with epidural catheters because IUCs might have been used for the patients. Finally, we did not assess other adverse events such as urethral trauma that might cause severe complications such as urosepsis and AKI [40] because we did not record these urethral traumata

Abbreviations: CI, confidence interval; ENT, ear, nose, and throat; IPW, inverse propensity-weighted.

| Surgery type      | General surgery\(a\) | ENT surgery\(b\) | Orthopedic surgery\(c\) | Others        |
|-------------------|----------------------|------------------|-------------------------|---------------|
| Referee           | 0.46                 | 0.28             | 0.83                    | 0.73–0.95     |
| Reference         | 0.73                 | 0.27             | 1.07                    | 0.97–1.19     |
| Hazard ratio      | 0.46                 | 0.28             | 0.83                    | 0.73–0.95     |
| \(95\%\) CI       | [0.42–0.50]           | [0.26–0.30]       | [0.73–0.95]             |               |
| \(p\)-value       | \(<0.001\)           | \(<0.001\)       | \(0.006\)               | \(0.001\)     |
| Hazard ratio      | Reference             | Reference         | Reference               | Reference     |
| \(95\%\) CI       | 0.73                 | 0.27             | 1.07                    | 0.97–1.19     |
| \(p\)-value       | \(<0.001\)           | \(<0.001\)       | \(0.006\)               | \(0.001\)     |

Notes.
Abbreviations: CI, confidence interval; ENT, ear, nose, and throat; IPW, inverse propensity-weighted.
\(a\) General surgery mainly consisted of cholecystectomy, hernia repair, and mastectomy.
\(b\) ENT surgery mainly consisted of septoplasty, tonsillectomy, tympanoplasty, and parotidectomy.
\(c\) Orthopedic surgery mainly consisted of treatment for upper or lower limb fracture, arthroplasty or osteotomy of the knee or foot, and spine surgery.

Our study demonstrated that intraoperative IUC use increases the risk of postoperative altered mental status, urinary tract infection, and prolonged hospital stay. However, the present results may be confounded by type 1 error because the altered mental status incidence was lower. According to our calculations, a RCT would require a total of 5184 patients (type 1 error \(< 0.05\) with a power of 0.8). Therefore, we conclude that patients undergoing minimally invasive surgery may not need IUCs during surgery, further studies are needed to evaluate the IUC use preoperatively.

Data availability
The data is sharing in a repository in which the authors deposited the data (University Hospital Medical Information Network Individual Case Data Repository [UMIN-ICDR]).

Ethical approval
The Jikei University School of Medicine Institutional Review Board approved the study (registration number: 31-030 [9529]; principal investigator: Toko Fukushima; date of registration: April 22, 2019).

Sources of funding
This work was supported by funding from our department. This research did not receive any specific grants from funding agencies in the public, commercial, or not-for-profit sectors.

Author contribution
1. Toko Fukushima, M.D. Contribution: This author helped design the study, collected and analyzed data, and prepared the manuscript.
2. Kazuhiro Shoji, M.D. Contribution: This author helped design the study, collected data, and prepared the manuscript.
3. Atsuko Tanaka, M.D. Contribution: This author helped design the study, collected data, and prepared the manuscript.
4. Yukari Aoyagi, M.D. Contribution: This author helped design the study, collected data, and prepared the manuscript.
5. Seiko Okui, M.D. Contribution: This author helped design the study, collected data, and prepared the manuscript.
6. Marie Sekiguchi, M.D. Contribution: This author helped design the study, collected data, and prepared the manuscript.
7. Ayako Shibai, M.D. Contribution: This author helped design the study, collected data, and prepared the manuscript.
8. Takanori Hiroe, M.D., Ph.D. Contribution: This author helped design the study, collected data, and prepared the manuscript.
design the study, analyzed the data, and prepared the manuscript.
9. Yasushi Mio, M.D. Contribution: This author prepared the manuscript.

Registration of research studies
1. Name of the registry: Indwelling catheters increase altered mental status and urinary tract infection risk: a retrospective cohort study.
2. Unique Identifying number or registration ID: UMIN000042664.
   (UMIN-CTR, University Hospital Medical Information Network – Clinical Trial Registry, https://upload.umin.ac.jp/cgi-open-bin/ctr/cgi?function=brows&action=browse&recptno=00048679&type=summary&language=j).
3. Hyperlink to your specific registration (must be publicly accessible and will be checked):
   https://upload.umin.ac.jp/cgi-bin/ctr/cgi-view_reg.cgi?recptno=13000498990.

Guarantor
Toko Fukushima is the guarantor of this study.

Consent
We waived the requirement for informed consent because the procedures did not require any additional intervention and because all patient data were anonymous.

Provenance and peer review
Not commissioned, externally peer-reviewed.

Declaration of competing interest
The authors disclose no conflicts of interest.

Acknowledgements
This work was supported by funding from our department. This research did not receive any specific grants from funding agencies in the public, commercial, or not-for-profit sectors.
We would like to thank our secretary and Editage (www.editage.com) for English language editing.

Appendix A. Supplementary data
Supplementary data to this article can be found online at https://doi.org/10.1016/j.amsu.2021.102186.

Fig. 3. Kaplan-Meier curves for postoperative hospital stay. (A) Kaplan-Meier curves for postoperative hospital stay censored at 30 d after surgery in unmatched surgical patients. The median observation time was 5 d (interquartile range, 2–7 d) for the control group and 7 d (interquartile range, 4–12 d; log-rank test, P < 0.001) for the indwelling urinary catheter (IUC) group. (B) Kaplan-Meier curves for postoperative hospital stay censored at 30 d after surgery in inverse propensity-weighted (IPW) surgical patients. The median observation time was 5 d (interquartile range, 2–11 d) for the control group and 7 d (interquartile range, 4–12 d; log-rank test, P < 0.001) for the IUC group.

Abbreviations: IUC, indwelling urinary catheter; IPW, inverse probability weighting.

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