Assessments of 14 Exploited Fish and Invertebrate Stocks in Chinese Waters Using the LBB Method

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Due to limited data availability, only a small subset of the exploited fish and invertebrate populations have been assessed along Chinese coasts, which precludes comprehensive management of the fisheries. Here, we applied a length-based Bayesian biomass estimator (LBB) to 14 fish and invertebrate stocks in China’s coastal waters to estimate their growth, length at first capture and current relative biomass ($B/B_0$, $B/B_{MSY}$) from length-frequency (LF) data. Of the 14 populations assessed, one have collapsed, nine are grossly over-exploited, and three are overfished. Moreover, 13 populations have smaller mean lengths at first capture ($L_c$) than the optimal length at first capture ($L_{c, opt}$), indicating that they are suffering from growth overfishing. Thus, larger mesh sizes in commercial fishery would increase both the catch and biomass for these species, given current levels of fishing mortality. Our results confirm that fishery resources in China’s coastal waters are strongly depleted, and that stricter management measures are needed to restore the abundance of China’s marine fisheries resources.

Keywords: stock assessment, data-poor stocks, length-frequency data, biomass estimation, LBB method

INTRODUCTION

As a result of decades of rapid development, China has become the country with the world’s largest fishery (FAO, 2018). Currently, China’s marine catch is around 10 million metric tons, with the bottom trawl fishery accounting for five million metric tons. However, this high reported catch is associated with very low catch per unit of effort, and the “fishing down” phenomenon was also shown to occur. Thus, to effectively rebuild fishery resources in China, robust management systems are needed, based on reliable stock assessment.

Most stock assessment models are designed to estimate fisheries and population-related reference points using fisheries-independent data sets, such as catch-at-age data and biomass estimates from scientific surveys (Methot and Wetzel, 2013). However, in China, due to limited data and expertise, only a small subset of the commercially exploited stocks, such as largehead hairtail Trichiurus lepturus and yellow croaker Larimichthys polyactis, have been assessed by dedicated stock assessment (Liu et al., 2013; Zhang and Chen, 2015). This situation is similar to that of many countries in the global South, where fishery research has been languishing.
In response to this issue, two basic approaches of methods were developed to perform stock assessments in data-sparse environments. One approach was to use catch time series and ancillary data to estimate Maximum Sustainable Yield (MSY) and related statistics (Schaef er, 1954; Froese et al., 2017). The other approach uses length-frequency (LF) data for inferences on the growth, mortality and hence populations’ response to fishing; see contributions in Pauly and Morgan (1987) and Pauly (1998), and Liang and Pauly (2017b) for a recent application to China’s coastal fisheries. The latter approach has the advantage that LF data is generally easier and cheaper to acquire, which makes LF-based methods the preferred approach in many situations.

LBB is a newly developed length-based Bayesian biomass estimation method (Froese et al., 2018) requiring LF data that are representative of the fishery under study. It then uses a Bayesian Monte Carlo Markov Chain (MCMC) to estimate growth and mortality parameters and relative stock size. Thus, even in data-poor situations, LBB assessment can be used directly by fishery managers. Also, LB results can be used as priors for stock assessment methods requiring an independent estimate of biomass relative to unfished biomass as input.

In this contribution, we applied LBB method to 14 fish and invertebrate populations in Chinese coastal waters to explore the extent of biomass depletion caused by fisheries, and to provide evidence for potential alternative harvest policies.

MATERIALS AND METHODS

Data Source

The data source and basic information of the 14 fish and invertebrate populations covered here are given in Supplementary Tables S1, S2 (see Supplementary Materials). Most LF data were read off from scientific papers.

However, the LF data for spiny red gurnard Chelidonichthys spinosus, hakodate sand shrimp Crangon affinis, red tonguessole Cynoglossus joyneri, slender lizardfish Saurida elongata, southern rough shrimp Trachysalambria curviostris were obtained from cruises performed in 2016 to 2017 on board of hired private fishing vessels operating in the Bohai Sea and North Yellow Sea, and which used the same gear as when fishing commercially.

General Description of LBB Method

The core of the LBB method is the von Bertalanffy growth function (VBGF; von Bertalanffy, 1938; Pauly, 1998), used to depict the growth in body length.

\[ L_t = L_{\text{inf}} \left[ 1 - e^{-K(t-t_0)} \right] \]  

where \( L_t \) is the length at age \( t \), \( L_{\text{inf}} \) is the asymptotic length, i.e., the mean length that the individuals of the species and stock in question would reach if they were to grow indefinitely, \( K \) is the rate by which \( L_{\text{inf}} \) is approached (year\(^{-1}\)), and \( t_0 \) is the age the fish would have at zero length if they always grew according to the VBGF.

Most species exploited by commercial fishery grow throughout their lives, and thus would approach \( L_{\text{inf}} \) if it were not for mortality, which is expressed by:

\[ N_{\text{f2}} = N_{t1} \cdot \exp - (Z \cdot (t_2-t_1)) \]  

where \( N_{t1} \) and \( N_{\text{f2}} \) are the numbers of a given cohort or a population at time 1 and 2, and \( Z \) is the instantaneous rate of total mortality, consisting of natural and fishing mortality, i.e., \( Z = M + F \) (Beverton and Holt, 1957; Pauly, 1998; Sparre and Venema, 1998).

Fishing gears have characteristic selection curves; the selection assumed in LBB is of the trawl type, i.e., very small individuals (<\( L_{\text{start}} \)) are not caught, all individuals past a certain size (>\( L_{\text{start}} \)) are caught, while the faction caught between \( L_{\text{start}} \) and \( L_{\text{end}} \) is an increasing function of the sizes. Such gear selectivity can be represented by

\[ S_L = \frac{1}{1 + e^{-a(t-L_c)}} \]  

where \( S_L \) is the fraction of individuals that are retained by the gear at length \( L \). In other words, \( S_L \) equals 0 when the length of individuals is less than \( L_c \), and is 1 when individuals exceed the length where full selection occurs. When individuals are subject to partial selection, \( S_L \) ranges from 0 to 1. In Eq. 3, \( a \) is the steepness of the ogive describing the length-dependent selectivity of the gear (Sparre and Venema, 1998). Thus, one can calculate the mean size at first capture (\( L_c \)), i.e., the length at which 50% of the individuals encounter the gear will be retained by it.

Combining Eqs. (1–3) and rearranging lead to:

\[ N_{L_i} = N_{L_{i-1}} \left( \frac{L_{\text{inf}} - L_i}{L_{\text{inf}} - L_{i-1}} \right)^{\frac{M}{M+K} S_{L_i}} \]  

and

\[ C_{L_i} = N_{L_i} S_{L_i} \]  

where \( N_{L_i} \) is the number of individuals at length \( L_i \), \( N_{L_{i-1}} \) refers to the number at the previous length \( L_{i-1} \). \( C \) is the number of individuals vulnerable to the gear, and all other parameters are as defined above. To minimize the parameter requirements of this approach method, the ratios \( M/K \) and \( F/M \) are output, instead of the absolute values of \( F \), \( M \), and \( K \); note that \( F/M = (F/K)/(M/K) \).

Also note that LBB, although it refers to “mean length at first capture” (\( L_c \)) in fact accounts for gradual selection as described by Eq. 3, including accounting for the fish that are caught at very small sizes (larger than \( L_{\text{start}} \), but below \( L_c \)) which are not compensated for by the larger fish that are not caught above \( L_c \), but below \( L_{\text{start}} \) (Silvestre et al., 1991).

When a species has more than one year LF data, the catch in numbers are made comparable between years by dividing both sides of Eq. 5 by their respective sums:
Then $M/K$ and $F/K$ can be deduced by fitting Eq. 4 to LF data.

Relative yield-per-recruit ($Y'/R$), as defined by Beverton and Holt (1966) can be computed, as presented by Froese et al. (2018) from:

$$
Y' = \frac{F}{M} \left( 1 - \frac{L_c}{L_{inf}} \right) \left( 1 - \frac{3 \left( 1 - \frac{L_c}{L_{inf}} \right)}{1 + \frac{M}{M/K + F/K}} \right) + \frac{3 \left( 1 - \frac{L_c}{L_{inf}} \right)^2}{1 + \frac{2}{M/K + F/K}} \left( 1 - \frac{L_c}{L_{inf}} \right) \left( 1 - \frac{3 \left( 1 - \frac{L_c}{L_{inf}} \right)}{1 + \frac{3}{M/K + F/K}} \right)
$$

(7)

Assuming CPUE proportional to biomass, dividing Eq. 7 by $F/M$ gives:

$$
CPUE' = \left( \frac{Y'}{R} \right) / \left( \frac{F}{M} \right) = \left( \frac{1}{1 + \frac{F}{M}} \right) \left( 1 - \frac{L_c}{L_{inf}} \right)^2 \left( 1 - \frac{3 \left( 1 - \frac{L_c}{L_{inf}} \right)}{1 + \frac{1}{M/K + F/K}} + \frac{3 \left( 1 - \frac{L_c}{L_{inf}} \right)^2}{1 + \frac{2}{M/K + F/K}} \right) - \frac{3 \left( 1 - \frac{L_c}{L_{inf}} \right)^3}{1 + \frac{3}{M/K + F/K}}
$$

(8)

Relative biomass of fish ($>L_c$) when $F = 0$ is then given by

$$
B_0 = \frac{L_c}{L_{inf}} \left( 1 - \frac{L_c}{L_{inf}} \right) \left( 1 - \frac{3 \left( 1 - \frac{L_c}{L_{inf}} \right)}{1 + \frac{1}{M/K}} + \frac{3 \left( 1 - \frac{L_c}{L_{inf}} \right)^2}{1 + \frac{2}{M/K}} - \frac{3 \left( 1 - \frac{L_c}{L_{inf}} \right)^3}{1 + \frac{3}{M/K}} \right)
$$

(9)
Palomares et al. (2018) was used to convert estimates of stock size ($B/B_0$) to biomass ($B$), which allows defining:

\[ \frac{B}{B_0} = \left( \frac{CPUE}{R} \right) \left( \frac{B_0 > L_c}{R} \right) \]  

(Froese et al., 2018).

Also, we have

\[ L_{opt} = L_{inf} \cdot \frac{3}{(3 + M/K)} \]  

where $L_{opt}$ is the length when a cohort of fish has its maximum biomass (Holt, 1958). This allows defining:

\[ L_{c, opt} = \frac{L_{inf} \cdot \left( \frac{2 + \frac{F}{M}}{1 + \frac{F}{M}} \right)}{\left( \frac{3}{3 + \frac{F}{M}} \right)} \]  

i.e., the mean length at first capture which maximizes the catch and the biomass for a given pair of $F/M$ and $M/K$ ratios.

A proxy for the relative biomass that can produce MSY ($B_{MSY}/B_0$) was obtained by re-running Eqs. 7–10 with $F/M$ = 1 and $L_c = L_{c, opt}$. With these parameters, current relative stock size ($B/B_{MSY}$) can be deduced; the nomenclature in Palomares et al. (2018) was used to convert estimates of $B/B_{MSY}$ into qualifiers of fisheries status.

If reliable estimates of $L_{inf}$ from independent sources are available, they can be used as priors in LBB analyses to decrease their uncertainty. In this contribution, we use $L_{inf}$ information from FishBase for fishes\(^1\) or SeaLifeBase for invertebrates\(^2\) as priors.

We used the 32.0 version of the LBB software (a package in R), which automatically generate priors for $L_c$ using the mean of $L_{10}$ and $L_{90}$, where $L_{10}$ and $L_{90}$ are lengths at 10 and 90% of the range of the LF data, respectively. $M/K$ is set as a default value 1.5; $Z/K$ prior is estimated according to Beverton and Holt [1957; $Z/K = (L_{inf} - L_{mean})/(L_{mean} - L_c)$]; $F/K$ prior equals to $Z/K-M/K$.

**RESULTS**

Figure 1 presents the LBB analyses for the five species with original LF data obtained from field surveys conducted in the Bohai Sea and North Yellow Sea 2016–2017 (see Supplementary Tables S1, S2). As might be seen, the LBB model fit to all five species, yielding estimates of $B/B_{MSY}$ ranging from 0.70 to 0.11 (Table 1). Among these five stocks, that of spiny red

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### TABLE 1 | Summary of LBB results for 14 populations.

| Scientific name (common name) | $L_{inf}$ (mm)* | $L_c/L_{c, opt}$ | $Z/K$* | $B/B_0$* | $B/B_{MSY}$* | Status\(^1\) |
|------------------------------|----------------|-----------------|--------|----------|------------|-------------|
| Chelidonichthys spinosus (spiny red gurnard) | 330 (326–336) | 0.62 | 9.10 (8.52–9.74) | 0.042 (0.029–0.059) | 0.11 (0.078–0.16) | Collapsed |
| Crangon affinis (hakodate sand shrimp) | 34.5 (34.0–35.1) | 0.46 | 3.28 (3.11–3.47) | 0.21 (0.16–0.29) | 0.58 (0.44–0.80) | Overfished |
| Cynoglossus joyneri (red tonguesole) | 274 (270–280) | 0.69 | 5.87 (5.52–6.3) | 0.10 (0.071–0.14) | 0.28 (0.19–0.39) | Grossly overfished |
| Saurida elongata (slender lizardfish) | 420 (411–426) | 0.78 | 6.32 (5.82–6.69) | 0.13 (0.099–0.17) | 0.37 (0.28–0.47) | Grossly overfished |
| Trachysalambria curvostris (southern rough shrimp) | 41.8 (41.1–42.5) | 0.78 | 2.92 (2.68–3.26) | 0.26 (0.15–0.39) | 0.70 (0.41–1.00) | Overfished |
| Hexagrammos agrammus (spotty-bellied greenling) | 284 (279–290) | 0.64 | 5.48 (5.09–5.93) | 0.076 (0.046–0.12) | 0.21 (0.12–0.30) | Grossly overfished |
| Muraenesox cinereus (Tanaka’s threadfin lizardfish) | 534 (525–541) | 0.67 | 4.91 (4.52–5.27) | 0.13 (0.085–0.19) | 0.35 (0.23–0.52) | Grossly overfished |
| Pennahia pawak (pawak croaker) | 242 (239–245) | 1.10 | 2.09 (1.93–2.3) | 0.70 (0.34–1.80) | 1.90 (0.094–4.90) | Healthy |
| Setipinna taty (scaly hairfin anchovy) | 200 (196–204) | 0.60 | 3.48 (3.27–3.7) | 0.19 (0.11–0.29) | 0.51 (0.30–0.79) | Overfished (1998) |
| Decapterus maruadsi (Japanese scad) | 279 (276–282) | 0.90 | 2.72 (2.49–2.95) | 0.30 (0.14–0.50) | 0.78 (0.36–1.30) | Overfished (2009) |
| Eynnis cardinals (threadfin porgy) | 278 (274–282) | 0.81 | 5.14 (4.81–5.52) | 0.17 (0.11–0.24) | 0.46 (0.30–0.68) | Grossly overfished (2009) |
| Liparis tanaka (Tanaka’s snallish) | 219 (215–221) | 0.72 | 3.39 (3.19–3.72) | 0.17 (0.096–0.26) | 0.46 (0.25–0.69) | Grossly overfished (2002) |
| Nemipterus bathybis (yellowbelly threadfin bream) | 545 (538–555) | 1.00 | 4.11 (3.80–4.49) | 0.29 (0.20–0.41) | 0.80 (0.55–1.10) | Slightly overfished (2005) |
| Prionchus macracanthus (red bigeye) | 531 (525–539) | 0.88 | 5.53 (5.13–6.03) | 0.11 (0.077–0.15) | 0.28 (0.20–0.40) | Grossly overfished (2010) |
| Trachysalambria curvostris (southern rough shrimp) | 244 (240–248) | 0.82 | 3.97 (3.77–4.32) | 0.24 (0.17–0.32) | 0.66 (0.46–0.88) | Grossly overfished (2005) |
| Hexagrammos agrammus (spotty-bellied greenling) | 239 (235–242) | 0.65 | 5.06 (4.86–5.33) | 0.10 (0.073–0.15) | 0.28 (0.19–0.40) | Grossly overfished (2009) |
| Prionchus macracanthus (red bigeye) | 307 (302–312) | 0.53 | 3.06 (2.91–3.32) | 0.33 (0.20–0.52) | 0.91 (0.55–1.40) | Slightly overfished (1999) |
| Pennahia pawak (pawak croaker) | 303 (297–308) | 0.50 | 3.70 (3.46–3.96) | 0.18 (0.12–0.25) | 0.48 (0.32–0.67) | Grossly overfished (2015) |

*The number between brackets stand for 95% confidence intervals for the parameter estimates in provides; † The number between brackets represent different years when the LF data were sampled.

\(^1\)www.fishbase.org

\(^2\)www.sealifebase.org
FIGURE 2 Application of the LBB method to 4 species of commercially exploited fish along the coast of China and for which length-frequency data were obtained from Zhou and Xu (2007), Xiong et al. (2009), Yan et al. (2011), and Ji (2014); see also Table 1 and Supplementary Tables S1, S2.

gurnard (*C. spinosus*) has collapsed, and the other four stocks are overfished or grossly overfished given the definitions of stock status in Palomares et al. (2018).

**Figure 2** presents similar results for four species with previously published LF data (see Supplementary Tables S1, S2). Here again, the LBB gives a good fit, and estimates of $B/B_{MSY}$ ranging from 1.90 to 0.21 (Table 1). Except for pawak croaker (*Pennahia pawak*), which appeared to be in a healthy state based on our estimates, the other stocks appear to be over-exploited.

**Figure 3** presents LBB applications to five species each, and for which LF data were available for two distinct periods (see also Supplementary Tables S1, S2). As might be seen, for all five species, the estimates of $L_{inf}$, $L_c/L_{c,opt}$ and $Z/K$ increased (Figure 3 and Table 1), while $B/B_{MSY}$ decreased (Table 1), as would occur if fishing effort increased over time.

Besides pawak croaker, 13 stocks have $L_c/L_{c,opt}$ less than one, implying that they are suffering from growth overfishing; thus increasing the mesh size of the gears exploiting them would lead to increases of their biomass and catch.

**DISCUSSION**

Full stock assessment generally requires fisheries-independent data sets, thus constrains its application in data-poor circumstances, which is the case for most China's coastal stocks (Geng et al., 2018). New computer-based methods came to the rescue with the advent of personal computers. One of the methods which the availability of computers made possible is the CMSY method (Froese et al., 2017), which uses a time series of catch and - if available - proxies of stock abundance (such as CPUE data) to infer MSY and fisheries reference points. However, in China, due to the fisheries catch misreporting, reliable catch data cannot be easily accessed (Jacquet et al., 2010).

The size composition (LF data) of exploited stocks has long been used to estimate exploitation rate and stock status (Beverton and Holt, 1957; Pauly and Morgan, 1987). Methods based on LF data can be either used directly in data-sparse fisheries management, or for providing prior of current relative biomass in other assessment models (Froese et al., 2017). LBB is a new Bayesian MCMC approach for the analysis of LF data from commercial catch.

The LBB method assumes constant recruitment, growth and mortality; thus it should not be used if these assumptions are strongly violated. The LBB method will also generate biased estimates if LF data are not representative, such as data sampled by gears that have a different selectivity or catchability than commercial gears, or in areas where only juveniles or adults occur. However, when reliable LF data are available, the LBB method can offer a window into the dynamics of the stocks, especially in data-poor situations.

In this contribution, we applied the LBB method to LF data of 14 fish and invertebrate populations from China's coastal waters. Among these, and given the nomenclature in Palomares et al. (2018), one population has collapsed and nine are grossly overfished. Our results are generally consistent with what is known of China's coastal resources, that is, the biomass of most populations are severely depleted, especially for main commercial species. The decline of fishery resources
FIGURE 3 | Application of the LBB method to 5 species of commercially exploited fish along the coast of China (Japanese scad Decapterus maruadsi, threadfin porgy Evynnis cardinalis, Tanaka’s snailfish Liparis tanakei, yellowbelly threadfin bream Nemipterus bathybius, and red bigeye Priacanthus macracanthus) and for with length-frequency data were obtained from two different time periods from Geng et al. (2018), Chen et al. (2012, 2013), and Zhang et al. (2016a,b); see Table 1 and Supplementary Tables S1, S2.
is confirmed by the decreased CPUE along Chinese coastal waters (Kang et al., 2018), the increased proportion of “trash fish” (Greenpeace, 2017; Zhang et al., 2020), and the occurrence of the “fishing down” effect in China’s coastal waters (Liang and Pauly, 2017a).

Despite the depletion of its coastal resources, China’s reported marine catch remains at a high level, which might be attributed to the illegal use of very small mesh size (~10 mm) in commercial bottom trawl fisheries along Chinese coastal waters. Based on our results, we concluded that among 14 populations in question, 13 stocks had smaller mean sizes at first capture than \( L_{c_{\text{opt}}} \), indicating that an increase in mesh size would increase both the catch from, and the biomass of these populations. This is in accordance with the finding of Liang and Pauly (2017b) and of Zhai and Pauly (2019).

Greenpeace (2017) released a report on China’s “trash fish” (fish that are too small for direct human consumption) fisheries, and pointed out that fishes in the larval or fingerling stages accounted for a large proportion of the catch of China’s coastal fisheries. To promote the restoration and sustainable use of marine fishery resources, China issued an announcement to regulate minimum catchable standards and juvenile fish proportion for 14 important commercial fish species, including Trichiurus japonicus, L. polyactis and Pampus argenteus (Ministry of Agriculture and Rural Affairs of the People’s Republic of China, 2018). This shows that China has taken this issue into account; how this regulation will be implemented will be crucial.

Our results show that reducing fishing mortality and increasing mesh size would benefit catch and biomass. This was backed up by Costello et al. (2016), who demonstrated that applying sound management reforms to Chinese fisheries could generate considerable increases in catch, biomass and profit relative to business as usual. In China, seafood is regarded as an excellent source of high-quality protein. The domestic consumption for seafood in China has significantly increased, from 7 kg per person in 1985 to around 25 kg per person in the early 2000s (Gao and Gao, 2005). The large demand for seafood has imposed enormous pressure on China’s coastal fisheries. However, it is obvious that the current management measures are not sufficient, and more stringent measures are needed to restore thriving coastal fisheries.

**DATA AVAILABILITY STATEMENT**

All datasets generated for this study are included in the article/Supplementary Material.

**ETHICS STATEMENT**

Ethical review and approval was not required for the animal study because our manuscript was based on survey cruise data or published data, and no live vertebrates or higher invertebrates were involved.

**AUTHOR CONTRIBUTIONS**

CL, WX, and SL were responsible for data collecting, formal analysis and writing original manuscript. DP contributed to the methodology, reviewing and editing the draft.

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**SUPPLEMENTARY MATERIAL**

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fmars.2020.00314/full#supplementary-material

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