Systematic Review and Meta-analysis: Use of Statins Is Associated with a Reduced Incidence of Oesophageal Adenocarcinoma

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Abstract

Purpose Laboratory studies have suggested that statins may have useful anti-cancer effects against Barrett’s epithelial cancer lines. A variety of effects have been reported in clinical studies.

Methods We performed a systematic review and meta-analysis of the association between statin use and the development of oesophageal cancer. Multiple databases were searched for studies reporting the association of statin use and oesophageal cancer. Meta-analysis on the relationship between statin use and cancer incidence was performed.

Results Twenty publications met eligibility criteria, yielding 22 datasets for meta-analysis. All were observational studies. Population-level studies included 372,206 cancer cases and 6,086,906 controls. Studies examining adenocarcinoma development in Barrett’s oesophagus included 1057 cancers and 17,741 controls. In patients with Barrett’s oesophagus, statin use was associated with a reduced incidence of adenocarcinoma (pooled adjusted odds ratio (OR) 0.59 (95% confidence intervals 0.50–0.68)), with no heterogeneity between 11 studies. Population-based studies demonstrated more heterogeneity but showed that statin use was associated with a lower incidence of both oesophageal adenocarcinoma (OR 0.57 (0.43–0.76)) and all oesophageal cancers (OR 0.82 (0.7–0.88)). Information on statin type, dose, and duration was reported too infrequently for statistical analysis but individual studies showed a tendency to a dose- and duration-dependant decrease in cancer incidence.

Conclusions Statin use is associated with a significantly lower incidence of oesophageal adenocarcinoma. This is seen in both Barrett’s cohorts and general populations. Further studies should focus on drug, dose, and duration and the interaction with other risk and preventative factors.

Keywords Barrett’s oesophagus · Chemoprevention · Cancer risk · Oesophageal carcinoma · Statins

Introduction

Oesophageal carcinoma (OC) affects 450,000 people worldwide and is the sixth leading cause of cancer-related mortality in the world [1]. The two most common histological subtypes, squamous cell carcinoma (OSC) and adenocarcinoma (OAC), together make up approximately 98% of this number. Global epidemiological studies show that whilst global incidence of OSC is greater than OAC, the incidence of OAC is now greater than OSC in the
developed countries of the West such as the UK, the Netherlands, and the USA [2, 3]. Currently, prognosis for oesophageal cancer remains one of the poorest amongst all cancers, with 5-year survival rates around 15% [3, 4]. This is largely due to the proportion of patients presenting at advanced disease stages. In OAC in particular, this has led to management strategies revolving around monitoring risk factors, most important being Barrett’s oesophagus (BO). To this effect, endoscopic surveillance as well as ablative procedure, principally radiofrequency ablation (RFA), has been developed [5]. Although surveillance and intervention strategies do seem to reduce the incidence of OAC, RFA is not without costs and complications [5].

However, due to the generally asymptomatic nature of BO, most cases of OAC arise outside surveillance programs. This is an area where chemopreventative strategies could make a difference to the general population. In the past 10 years, experimental evidence has shown that statins [hydroxymethylglutaryl-CoA reductase inhibitors (HMG-CoA inhibitors)] could potentially play a role in reducing progression of BO to OAC [6, 7]. The mechanisms proposed include upregulation of pro-apoptotic proteins such Bad and Bax, as well as reduced cell proliferation and enhanced apoptosis through reduced activation of signalling G-proteins as a consequence of reduced levels of melatonin due to inhibition of the HMG-CoA reductase [6, 7]. Through these mechanisms and possibly other mechanisms, statins attenuate the malignant behaviour of OAC cell lines [6–10]. Evidence from a variety of clinical studies supports this hypothesis. Four years ago, two separate meta-analyses using slightly different methods came to almost identical conclusions: statin use was associated with a significant 43% reduction in the incidence of OAC in Barrett’s cohorts and a smaller [19%] reduction in the incidence of all oesophageal cancers in population studies [11, 12]. However, these studies highlighted areas of lack of information including the precision of the inverse association with statin use, effects of confounding risk factors, and effects in different populations. Since the publication of these meta-analyses, multiple additional studies have been published and we have performed an updated systematic review and meta-analysis examining the association between statin use and OAC with the aim of further refining the association and in particular defining knowledge gaps that could be beneficially be the target of future studies [12]. The inclusion of more recent data has enabled us to specifically examine the association between statin use and oesophageal cancer in three categories: statin use in relation to malignant progression to OAC in Barrett’s oesophagus cohorts, the association between statin usage and OAC in a population cohorts, and the association between statin usage and all oesophageal cancers on a population scale.

Methods

This systematic review was performed according to the Preferred Reporting Items for Systematic reviews and Meta-Analyses Guidelines (PRISMA guidelines) and as previously described [12, 13].

Selection Criteria

Eligible study design characteristics included randomised controlled trials, observational studies (cohort and case-control design) that met the following inclusion criteria: statin exposure status with reliable evidence through record linkage or otherwise, reported disease incidence (oesophageal adenocarcinoma or high-grade dysplasia, or oesophageal cancer), and provision of hazard ratio (HR), odds ratio (OR), or relative risk (RR), and in the case of their absence, provision of data for their calculation.

Data Sources and Search Strategy

A systematic literature search of Medline, Embase, Web of Science, Cochrane database, and Google Scholar from inception until 4th February 2017 was performed in order to identify all relevant studies that investigated oesophageal cancer incidence in statin users. A combination of keywords and medical subject heading terms were used including “oesophageal neoplasm”, “Barrett’s oesophagus”, “oesophageal adenocarcinoma”, “statin”, and “hydroxymethylglutaryl-CoA reductase inhibitor”. Two investigators (Thomas T, Beales ILP) then independently reviewed and excluded articles that were not relevant to the research topic. A manual search of reference lists as well as relevant review articles was undertaken to screen for potential articles.

Data Extraction and Quality Assessment

Data was extracted from the selected studies into a standardised table by the two chief investigators. This table is provided in Table 1. Data was organised into the three subgroups being examined. Information regarding duration and dose of statin use was also extracted alongside number of statin/non-statin users to further assess the impact of statins. Unadjusted and adjusted risk ratios were also extracted where available. We checked the validity of the included studies based on possibility of confounding and potential for misclassification of tumour pathology and/or drug exposure. Risk of bias assessment was focused on the selection of participants, comparability of cases and controls (with any adjustments for confounding), and methods used in ascertaining drug exposure and outcomes. The quality of all studies was assessed using the Newcastle-Ottawa Quality Assessment Scales by two investigators independently using the star-rated system.
| First author [ref] | Location, setting | Study design | Total number of patients | Primary outcome of interest | Patient on statins | Patient not on statins | Variables adjusted for | Newcastle-Ottawa score |
|-------------------|-------------------|--------------|--------------------------|----------------------------|-------------------|-----------------------|------------------------|-----------------------|
| Nguyen 2010 [14]  | US; multicentre, hospital-based | Retrospective cohort | 344 patients | OAC 33 | Incidence of OAC in Barrett’s cohort | 6 | 63 | 27 | 188 | 1, 2, 7, 9 | 8a |
| Nguyen 2009 [15]  | US; multicentre, hospital-based | Nested case control | 812 patients | OAC 116, BO 696 | Incidence of OAC in Barrett’s cohort | 41 | 336 | 75 | 360 | 1, 2, 3 | 9a |
| Kastelein [16]     | The Netherlands; multicentre, hospital-based | Prospective cohort | 570 patients | HGD/OAC 38 | Incidence of OACC in Barrett’s cohort | 9 | 200 | 29 | 332 | 1, 2, 7, 8, 9 | 9a |
| Kantor [17]        | US; hospital-based | Prospective cohort | 411 patients | OAC 56 | Incidence of OAC in Barrett’s cohort | 6 | 50 | 50 | 305 | 1, 2, 4, 9 | 9a |
| Beales 2012a [18]  | UK; hospital-based | Case control | 255 patients | OAC 85, BO 170 | Incidence of OAC in Barrett’s cohort | 17 | 60 | 68 | 110 | 1, 2, 4, 5, 6, 8 | 8a |
| Krishnamoorthi [19]| US; population-based | Cohort | 9660 patients | OAC 103 | Incidence of OAC in Barrett’s cohort | NR | NR | NR | NR | 1, 2, 4, 6, 9 | 8a |
| Iyer [20]          | UK; population-based | Cohort | NR | Incidence of OAC in Barrett’s cohort | NR | NR | NR | NR | 6a |
| Agrawal [21]       | US; hospital-based | Case control | 583 patients | Incidence of OAC in Barrett’s cohort | 55 | 307 | 60 | 161 | 1, 4, 5, 6, 9 | 7a |
| Masclee [22]       | UK and the Netherlands; multicentre, population-based | Nested case control | 777 patients | OAC 45, BO 732 | Incidence of OAC in Barrett’s cohort | 12 | 253 | 33 | 479 | None listed | 7a |
| Cooper [23]        | UK; population-based | Nested case control | 3749 patients | OAC 55 | Incidence of OAC in Barrett’s cohort | 18 | 1124 | 37 | 2570 | 1, 2, 4 | 8a |
| Nguyen 2015 [24]   | US; population-based | Nested case control | 1167 patients | OAC 311, BO 856 | Incidence of OAC in Barrett’s cohort | 125 | 462 | 186 | 394 | 1, 4, 6, 9 | 9a |
| Kaye [25]          | UK; population-based | Case control | 530 patients | OAC 100, control | Incidence of any OC in | 9 | 34 | 91 | 396 | 1, 2, 4, 6 | 8a |
| First author [ref] | Location, setting | Study design | Total number of patients | Primary outcome of interest | Patient on statins | Patient not on statins | Variables adjusted for | Newcastle-Ottawa score |
|-------------------|-------------------|--------------|--------------------------|----------------------------|-------------------|-----------------------|-----------------------|-------------------------|
| Friedman [26]     | US; population-based | Cohort      | 4,413,100 patients       | Incidence of any OC in population | 68 762           | 361,802 4,050,468  | NR 9*                  |                         |
| Hippiseley-Cox [27] | Male UK; population-based | Cohort      | 989,729, OC 1225         | Incidence of any OC in population | 216 120,866 1009 867,638 | 1, 4, 6 9* |                         |                         |
|                  | Female UK; population-based | Cohort      | 1,013,565, OC 584        | Incidence of any OC in population | 78 104,670 506 908,311 | 1, 4, 6 9* |                         |                         |
| Vinogradava [28]  | UK; population-based | Nested case control | 16,200 patients OC 3159, control 13,041 | Incidence of any OC in population | 496 2106 2663 10,935 | 4, 6 9* |                         |                         |
| Bhutta [29] (abstract) | UK; population-based | Prospective cohort | 18,484 patients OAC 581, OSC 322 | Incidence of any OC in population | 615 2539 3101 12,229 | NR 7* |                         |                         |
| Lai [30]          | Taiwan; population-based | Case control | 2745 patients OC 549, control 2196 | Incidence of any OC in population | 49 238 500 1958 | 1, 2, 8, 9 6* |                         |                         |
| Chan [31]         | Taiwan; population-based | Case control | 985 patients OC 197, control 788 | Incidence of any OC in population | 29 131 168 657 | 8, 9 7* |                         |                         |
| Alexandre [32]   | EAC UK; population-based | Nested case control | OAC 581 Control 2167 | Incidence of OAC in population | 60 222 521 1945 | 4, 6, 9 8* |                         |                         |
| EGJA              |                   | Case control | OAC 213 Control 783      | Incidence of OAC in population | 20 79 193 704 | 4, 6, 9 8* |                         |                         |
| Beales 2012b [33] | UK; hospital-based | Case control | 560 patients OAC 112, control 448 | Incidence of OAC in population | 19 158 93 290 | 1, 2, 4, 5, 6, 9 8* |                         |                         |

Variables adjusted for (1) age, (2) sex, (3) race, (4) smoking, (5) alcohol use, (6) obesity, (7) BO length, (8) oesophagitis or reflux symptoms, (9) medications (NSAIDs/aspirin)

NR not reported

* Statin use at baseline and not during course of study
as previously described [12, 34]. Any discrepancies in scoring were then resolved through combined reassessment and consensus by all authors. The final quality assessment scores are listed in Table 1.

**Outcomes Assessed**

The primary analysis examined the association of statin and the incidence of oesophageal carcinoma as well as oesophageal adenocarcinoma specifically, through comparison of users and non-users. Subgroup analyses were used to investigate this further. Three categories were devised according to study cohort characteristics. These included OAC incidence in a population cohort, OAC incidence in Barrett’s oesophagus cohort, and incidence of all oesophageal cancers in a population cohort. Secondary analysis was focussed on examining these studies for any potential duration or dose relationship between statin and cancer incidence. In order to perform this, we limited analyses on duration or dose to studies that explicitly provided reliable statin use description data.

**Statistical Analysis**

Review Manager (Revman) version 5.3 (Nordic Cochrane Center, Copenhagen, Denmark) was used to calculate the pooled risk ratio (compiling ORs or HR from individual studies) using the inverse variance method, random effects model as previously described [12]. Due to the relative rarity of outcomes, OR were considered as approximations of HR and RR. Analysis was carried out on unadjusted as well as adjusted risk ratios. Statistical heterogeneity was assessed using the Cochrane $I^2$ statistic, with $I^2 > 25\%$ indicating moderate statistical heterogeneity and $I^2 > 50\%$ indicating a substantial level of heterogeneity. A sensitivity analysis was performed by separately omitting one study at a time to assess if the pooled estimate had changed significantly compared to the results of all pooled studies.

**Results**

The search strategy identified 110 individual potentially suitable publications. Through the eligibility and screening processes, a total of 20 publications were included in the meta-analysis [14–33]. This totalled 22 different relevant studies as two publications reported separate data and adjusted odds ratios for different cohorts (either separate data for males and females or separate histological subtype of oesophageal cancer) [27, 32] (Fig. 1). All studies were published in English. The studies identified were clearly split into three categories based upon the specific population being examined: the association between statin use in the community and incidence of OAC, the relationship between statin use in a BO cohort on OAC rates, and the association with all oesophageal cancer on a population scale. Table 1 lists the studies included and their important characteristics. Two of the included studies were only published in abstract format; however, sufficient relevant information necessary for analysis was extracted from them [20, 29]. The other studies were in full peer-reviewed format.

For each planned subgroup analyses, separate meta-analyses were conducted on adjusted and unadjusted data. This was due to the heterogeneity in presentation of risk factor adjustment in the studies being analysed. As in the previous meta-analysis [12], heterogeneity in the reporting of statin dose and duration meant statistical analysis of this data would not be feasible.

**Characteristics of Included Studies**

Twenty-two datasets reported on a total of 373,263 cases and 6,105,704 controls. As a proportion, 3.7% of the total population were statin users. Of 236,608 statin users, approximately 0.82% developed HGD/OC. Of 6,231,642 non-statin users, 5.90% developed HGD/OC. Details of the included studies are given in Table 1. The studies were from a variety of geographic locations with five from the USA and seven from the UK. No randomised control trials were identified during the search and in total, data were available from eight cohort, six nested case-control, and six case-control studies. The majority of outcome assessments in nested case-control studies [23, 24, 32] and cohort study [19] were done through record linkage. Two case-control studies [18, 33] ascertained exposure through the use of structured interviewers where participants were not completely blind to case-control status. One case-control study ascertained exposure through the use of secure records (i.e. medical notes). Eleven studies [14–24] examined statin use and the incidence of OAC in Barrett’s cohort. Eight datasets reported [25–31] the impact of statins on incidence of any OC in population and two studies (with a total of 3 datasets [32, 33]) looked at the relationship between statins and OAC in the general population.

Table 2 displays the baseline characteristics of the population from each study. Mean age at BO diagnosis or age of control in nested case-control studies ranged from 60.7 to 70.9 years. In case-control studies, the mean age of diagnosis of OC ranged from 60.9 to 70.7 years. With the exception of one study [32], the proportion of men as part of study population exceeded 60%. Length of BO was only reported in two studies [15, 16]. Three studies reported rates of reflux symptoms and endoscopic oesophagitis [16, 22, 31]. Statin use in the included study populations ranged from 13 to 54%. Proton pump inhibitors (PPI) used in the study populations ranged from 51 to 94%.
Quality Assessment

The studies included in the meta-analysis were ranked as being medium (11) to high quality (11) with the exception of two studies [20, 30]. A more detailed breakdown of study quality assessment can be found in Table 1. Baseline characteristics in patients are relatively consistent across studies producing an appropriate pooled population for oesophageal carcinoma. This was a population consisting of a majority Caucasian male population of age greater than 60 years. Overall 13 studies included adjusted for concomitant potentially chemopreventative medication in the form of aspirin/non-steroidal anti-inflammatory drugs (NSAIDs). The majority of studies also adjusted for two other main risk factors involved in the development of OC: smoking and obesity. However, adjustment for other risk factors was variable between studies. Due to the variability in correcting for and reporting potential confounders, the pooled data for both un-adjusted and adjusted odds ratios were separately analysed.

Statin Use in Barrett’s Cohorts Progressing to Adenocarcinoma

A total of 11 studies were included within this analysis. This included five cohort studies, two case-control studies, and four-nested case-control studies. The total sample included a minimum of 1057 cancer/HGD cases and 17,741 controls with non-cancerous with BO (the actual numbers included in one study are not available [20]). All studies adjusted for age and gender except for one study that did not adjust for gender [24]. Three studies adjusted for race [14, 16, 18]. Six studies adjusted for smoking [17–19, 23, 24, 33]. Only one study adjusted for alcohol use [18]. Four studies adjusted for obesity [17–19, 24]. Two studies adjusted for length of BO [15, 16] and reported reflux with/without endoscopic oesophagitis [16, 18]. Five studies adjusted for concomitant medication use (specifically NSAID/aspirin use) [14–16, 18, 22]. One study [15] did not report adjusted OR in the final publication and hence was not included in the relevant meta-analysis.
### Table 2  Baseline characteristics of patients in studies included in the meta-analysis

**a. EAC incidence in Barrett’s patients**

| Study (reference) | Cohort: age at BO diagnosis (years, mean (SD)) | Gender (% men) | Race (%Caucasian) | Obesity (% with BMI > 30) | Smoking (% ever) | Reflux symptoms or endoscopic oesophagitis | Length of BO (>3 cm) | Other medication use | Other medication use |
|-------------------|---------------------------------|----------------|-------------------|------------------------|-----------------|------------------------------------------|---------------------|---------------------|---------------------|
| NGuyen 2009 [16] C | 60.7 (11.8)                      | 94.2           | 90.4              | NR                     | NR              | NR                                       | 29.1                | 49.1                | 25.3                |
| Nguyen 2010 [18] CC | 65                             | 97             | 74                | NR                     | NR              | NR                                       | 57.6                | 46.4                | 94.8                |
| Kastelein 2011 [20] C | 61 (53–68)                       | 71             | NR                | 19                     | 19 (current)    | 29/9                                     | 100                 | 20.4                | 18.9                |
| Kantor 2012 [21] C | NDR                             | 81.3           | NR                | NR                     | 64              | NR                                       | NR                  | 41.2 (current)      | 13.6 (current)      |
| Beales 2012a [22] CC | Case 68.9 (12.1) EAC Control 67.3 (12.0) BO | 80             | 100               | NR                     | 67.1            | 59.4                                     | N/A                 | 21.1                | 33                  |
| Krishnamoorthi 2016 [26] C | 63 (13.5)                       | 62.6           | NR                | 21.32                  | 51.9            | NR                                       | NR                  | 62.27               | 27.62               |
| Iyer 2014 [25] abstrac C | 63.04 (13.79)                   | 63.89          | NR                | NR                     | NR              | NR                                       | NR                  | 45                  | 88.3                |
| Agrawal 2014 [27] CC | Case 65.9 (±9.5) EAC Control 61.4 (±11.4) BO | 98             | 96                | 74                     | 64.8            | NR                                       | NR                  | 45                  | 88.3                |
| Masclee 2014 [28] nCC | UK 63 (SD 13.8) NL 61.2 (SD 13.4) | UK 63          | NR                | UK 51.7                | 0.04            | N/A                                      | Aspirin: 48.93  | 35.97               | 86.8                |
|                    |                                 | NL 62          |                   | NL 49.48               |                 | N/A                                      | NSAIDs: 0.06        | UK 16.33            | 51.7                |
| Cooper 2014 [29] nCC | 63 (52–72) BO                   | 63             | NR                | NR                     | 55              | NR                                       | NR                  | 1124 (30)           | 3569 (97)           |
| NGuyen 2015 [32] nCC | Case 64.7 (9.1) Control 64.5 (9.1) | 100            | 89.1              | 47                     | 19              | NR                                       | NR                  | 40.2                | 63                  |
|                    |                                 | 100            | 85.7              | 42.1                   | 13.4            | NR                                       | NR                  | 54                  | 80.4                |

**b. All oesophageal cancer incidence in population-based studies**

| Study | Cohort: age at BO diagnosis, years (mean (SD)) | Gender (% men) | Race (%Caucasian) | Obesity (% with BMI > 30) | Smoking (% ever) | Reflux symptoms or endoscopic oesophagitis | Length of BO (>3 cm) | Other medication use | Other medication use |
|-------|-----------------------------------------------|----------------|-------------------|------------------------|-----------------|------------------------------------------|---------------------|---------------------|---------------------|
| Kaye 2004 [31] CC | NR                                           | NR             | NR                | NR                     | NR              | NR                                       | NR                  | NR                  | NR                  |
| Friedman 2008 [15] C | NR                                           | NR             | NR                | NR                     | NR              | NR                                       | NR                  | NR                  | NR                  |
| Hippisley-Cox 2010 (men) [17] C | NR                                           | NR             | NR                | NR                     | NR              | NR                                       | NR                  | NR                  | NR                  |
| Hippisley-Cox 2010 (women) [17] C | NR                                           | NR             | NR                | NR                     | NR              | NR                                       | NR                  | NR                  | NR                  |
| Vinogradava 2011 [19] nCC | NR                                           | NR             | NR                | NR                     | NR              | NR                                       | NR                  | 16                  | NR                  |
| Bhutta 2011 [33] C | NR                                           | NR             | NR                | NR                     | NR              | NR                                       | NR                  | 16                  | NR                  |
| Lai 2012 CC 241 | Case 60.9 (12.9) | 92.9           | NR                | 0.55                   | 1.64            | NR                                       | NR                  | 34.1, 95.8          | 8.93                | 47.7                |
examining adjusted data. One study [19] used a time-varying model to assess HR. This was used to give a more reliable estimate, as it will take into account the time windows when the patients are taking or not taking statins. This study was incorporated into the adjusted OR model.

The pooled unadjusted data showed a significant inverse association between statin use and the incidence of OAC in Barrett’s cohorts (pooled OR 0.54 (95% CI 0.46 to 0.63)) ($P < 0.0001, I^2 = 0\%$) (Fig. 2). This confirmed the pooled adjusted ORs (OR 0.59 (95% CI 0.50 to 0.68)) ($P < 0.00001, I^2 = 0\%$) (Fig. 3). There was no significant heterogeneity in the studies and sensitivity analyses showed that the omission of any single study had no effect on the overall results. Both case-control and cohort studies produced very similar results and the results of the study only presented in abstract form for which less detail is available were consistent with all the other studies [20].

There was major variation in the reporting of duration and dose of statin use within the included studies. This outcome was therefore unsuitable for statistical analysis. Nguyen et al. [14] suggested that duration of statin use greater than 1 year was associated with a reduced incidence of OAC (aOR 0.52; 95% CI 0.3 to 0.91). Beales et al. [22] reported that statin use of greater than 5 years was associated with a reduced incidence of OAC (aOR 0.41; 95% CI 0.15 to 0.85). Nguyen et al. [24] also reinforced this general finding with both unadjusted and adjusted OR for the period of 6 to 18 months showing a reduced incidence of OAC (OR 0.41; 95% CI 0.15 to 0.85). Nguyen et al. [24] also reinforced this general finding with both unadjusted and adjusted OR for the period of 6 to 18 months showing a reduced incidence of OAC (OR 0.41; 95% CI 0.15 to 0.85). Nguyen et al. [24] also reinforced this general finding with both unadjusted and adjusted OR for the period of 6 to 18 months showing a reduced incidence of OAC (OR 0.41; 95% CI 0.15 to 0.85). Nguyen et al. [24] also reinforced this general finding with both unadjusted and adjusted OR for the period of 6 to 18 months showing a reduced incidence of OAC (OR 0.41; 95% CI 0.15 to 0.85).

In terms of dosage, Beales et al. [18] reported that statin dose above 40 mg equivalent of simvastatin was associated with a more pronounced inverse association than doses below 40 mg simvastatin dose equivalents (aOR 0.31; 95% CI 0.05 to 0.97 compared to 0.59; 95% CI 0.27 to 0.98, respectively). The other studies did not provide data on statin dose, type, or duration.

### Statin Use in Population Cohort Progressing to Adenocarcinoma

This subgroup analysis consisted of data extracted from two publications. One paper reported separate data from OAC and oesophagogastric junction adenocarcinoma (OGJA); due to the potential difference in biology, these were analysed and reported separately [32]. All studies were case-control studies. The total sample size was 4305 patients. This consisted of 907 cases of OAC with 3398 controls. All studies included adjusted and unadjusted data, and hence, both sets of data were compiled into the meta-analysis. Alexandre et al. [32] adjusted for risk factors such as smoking, obesity, and concomitant...
medication use in the form of NSAIDs, aspirin, and PPIs. Beales et al. [33] adjusted for age, gender, smoking, alcohol, obesity, and medications.

Pooled unadjusted ORs showed a non-significant association with wide confidence intervals with significant heterogeneity (0.73 (0.41–1.32)) (Fig. 4). However, the meta-analysis of pooled adjusted data (Fig. 5) showed a significant negative association between statin use and OAC incidence (OR 0.57; 95% CI 0.43 to 0.76) ($P < 0.001$), with no significant heterogeneity between studies ($I^2 = 0\%$).

When examining duration of statin use, Beales et al. [33] showed duration of use greater than 5 years had greater inverse association (adjusted OR 0.29; 95% CI 0.1 to 0.67) compared to other time periods such as 2 to 5 years (adjusted OR 0.51; 95% CI 0.19 to 0.97). Both studies [32, 33] examined dose with a common dose margin of 40 mg simvastatin. Beales et al. [33] reported that doses greater than 40 mg simvastatin were associated with a greater reduction in risk (adjusted OR 0.16; 95% CI 0.05 to 0.6) than doses less than 40 mg (adjusted OR 0.53; 95% CI 0.25 to 0.86). Although the data provided for statin dose less than 40 mg in Alexandre et al. [32] are not statistically significant, the statin dose for greater than 40 mg shows a substantial reduction in risk of OAC in a population cohort (adjusted OR 0.51; 95% CI 0.27 to 0.96).

### Statin use in Population Cohort Progressing to Oesophageal Cancer (All Types)

This subgroup analysis examined the potential effect of statins on the incidence of all types of oesophageal carcinomas in population cohorts. A total of eight cohorts were used in this analysis. This consisted of seven studies with one study [27] providing two populations consisting of the male and female population analysed separately. The sample size totalled 645,338 patients. There were 371,400 cases of oesophageal carcinomas in a background of 6,083,938 patients from population cohorts. Three studies adjusted for age [25, 27, 30], smoking [25, 27, 28], and obesity [25, 27, 28]. Two studies adjusted for gender [25, 30]. Lai et al. [30] and Chan et al. [31] both adjusted for symptomatic reflux and endoscopic oesophagitis along with concomitant use of medications (aspirin/NSAIDs). One study [26] did not provide details of any

#### Table 2

| Study or Subgroup | Weight | Odds Ratio IV, Random, 95% CI | Odds Ratio IV, Random, 95% CI |
|-------------------|--------|-------------------------------|-------------------------------|
| Agrawal 2014      | 14.5%  | 0.48 [0.32, 0.73]             |                               |
| Beales 2012a      | 6.5%   | 0.46 [0.25, 0.85]             |                               |
| Cooper 2014       | 7.8%   | 1.04 [0.59, 1.82]             |                               |
| Iyer 2014         | 8.9%   | 0.41 [0.24, 0.69]             |                               |
| Kantor 2012       | 4.9%   | 0.53 [0.26, 1.09]             |                               |
| Kastelein 2011    | 4.2%   | 0.52 [0.24, 1.11]             |                               |
| Masclee 2015      | 5.4%   | 0.69 [0.35, 1.36]             |                               |
| Nguyen 2009       | 2.9%   | 0.66 [0.26, 1.68]             |                               |
| Nguyen 2010       | 14.8%  | 0.59 [0.39, 0.88]             |                               |
| Nguyen 2015       | 30.2%  | 0.50 [0.37, 0.66]             |                               |

**Total (95% CI)** 100.0% 0.54 [0.46, 0.63]

Heterogeneity: Tau^2 = 0.00; Chi^2 = 7.96, df = 9 ($P = 0.54$); $I^2 = 0\%$

Test for overall effect: $Z = 7.64 (P < 0.00001)$

Fig. 2 Meta-analysis of pooled unadjusted odds ratios for the effect of statin use on the development of oesophageal adenocarcinoma in patients with Barrett’s oesophagus.

Fig. 3 Meta-analysis of pooled adjusted odds ratios for the effect of statin use on the development of oesophageal adenocarcinoma in patients with Barrett’s oesophagus.
adjustment in both study design and, statistically, hence was unable to be included in the adjusted OR meta-analysis.

The result of the meta-analysis of unadjusted odds ratios (1.07 (0.90–1.26)) (Fig. 6) was statistically insignificant. This was largely due to the influence of two studies [27, 25]. The unadjusted and adjusted ORs within the study by Hippiseley Cox et al. [27] were considerably different. There was substantial heterogeneity in the pooled unadjusted data (I² levels 82%).

The pooled analysis of adjusted ORs demonstrated that the use of statins in a population cohort was associated with a significantly lower incidence of all oesophagus (OR 0.82; 95% CI 0.76 to 0.88) (P < 0.00001) with no heterogeneity in the data (I² = 0%) (Fig. 7). Sensitivity testing showed that the significance of the result was not influence by the exclusion of any single study.

Similar to the previous analysis, there was significant heterogeneity in the methods of data collection and reporting regarding duration and dosage of statins. This rendered statistical analysis inappropriate. Friedman et al. [26] actually reported statins were associated with an increased risk of OC in the male population (adjusted OR 1.7; 95% CI 1.05 to 2.75).

However, with such large confidence intervals in place, this data are potentially unreliable. Data from Lai et al. [30] showed that statins were associated with a reduced incidence of OAC in a population cohort particularly when used for greater than 12 months (adjusted OR 0.14; 95% CI 0.04 to 0.56). In terms of statin doses, only Hippiseley-Cox et al. [27] provided detailed information. In males, doses of simvastatin equivalent to 40–80 mg showed an aOR 0.66; 95% CI 0.48 to 0.91 with the incidence of all oesophageal cancers. Doses of 10 mg atorvastatin were associated with reduced incidence of OC (aOR 0.68; 95% CI 0.49 to 0.95). In females, this trend was reversed with data suggesting dosages of 10 mg atorvastatin (aOR 0.67; 95% CI 0.49 to 0.95) were more effective in reducing OC rates in a population cohort than 40–80 g simvastatin (aOR 0.82; 95% CI 0.66 to 0.91).

**Discussion**

The primary objective of this meta-analysis was to examine association of statins with the development of oesophageal cancers, particularly oesophageal adenocarcinoma. Our pooled estimates show that statins have a strong association with a reduction in the risk of OAC (approximately 41% reduction of risk) particularly when used in patients with pre-existing Barrett’s oesophagus. The adjusted data from 11 studies provides a reasonably narrow confidence interval for this figure. These data are similar to those reported in the previous meta-analyses: Singh et al. [11] and Beales et al. [12] adjusted OR = 0.59; 95% CI, 0.45–0.78 and adjusted OR = 0.57; 95% CI, 0.43–0.75, respectively.

A similar degree of reduction in OAC was associated with statin use in population cohorts. The pooled adjusted odds ratios showed a reduction of 43% in OAC incidence. However, with only three datasets included, it is relatively the most underpowered of all analyses included in this study. However, this is a new finding; previous systematic reviews did not allow statistical analysis of this particular subgroup.

**Fig. 4** Meta-analysis of pooled unadjusted odds ratios for the effect of statin use on the development of oesophageal adenocarcinoma in population-based studies

**Fig. 5** Meta-analysis of pooled adjusted odds ratios for the effect of statin use on the development of oesophageal adenocarcinoma in population-based studies
Further studies are warranted to investigate this effect more thoroughly especially in light of the rising incidence rates of OAC. It is notable that this subgroup analysis included a specific oesophagogastroduodenal adenocarcinoma group, whilst there may be some important biological differences between this and more classical OAC [8], it was felt appropriate to include this cohort as the majority of the other studies included in this systematic review included junctional cancers within the cancer cases, making no distinction between the OAC and OGJA for statistical analysis. Two studies [18, 33] specifically excluded Siewert III OGJA cancer.

It remains an area of active research whether statins have beneficial anti-cancer effects merely on progression of non-dysplastic BO or whether they are actually associated with a reduction in the development of BO itself. Whilst most of the in vitro studies have examined the effects of statins on various models of Barrett’s oesophagus, recent data including a meta-analysis have reported that statin use is associated with a reduced incidence of Barrett’s oesophagus [35].

Statin use in the population is associated with a reduced incidence (by 18%) of all types of oesophageal cancers. Again, three of the major studies [27–29] included in the analysis provided very consistent data in describing this effect. The previous meta-analysis [12] suggested very similar reduction of risk at OR = 0.81; 95% CI 0.75–0.88. There was much more heterogeneity in the unadjusted results examining statin use at the population level and all oesophageal cancers than in the other analyses. The heterogeneity was significantly reduced when comparing the odds ratios adjusted for other variables. This probably reflects the wider variety of geographical and health care locations included in this group (certainly compared to the more homogeneous secondary care Barrett’s cohorts) and reinforces the problems with potential confounding by indications and risk factors in studies at that population level. Despite this, there were no obvious outlying studies and the pooled analysis of the adjusted odds ratios showed no significant heterogeneity.

Most of the published data have examined the relationship between statins and OAC; there are much less data concerning squamous cancer. Although there are well-performed laboratory studies providing biological plausibility and exploring the effects of statins on oesophageal adenocarcinoma and non-malignant Barrett’s cells [6, 7, 10], these corroborating studies are lacking as regards oesophageal squamous cancer. Given the contrasting biology and epidemiology between oesophageal adenocarcinoma and squamous cancers, it would not be surprising if the effect of statins differed between the two. The relatively smaller apparent protective effect against all oesophageal cancers combined compared to specifically against OAC may imply that statins have less effect on the...
development of squamous cancer but further studies are indicated. The relative and absolute numbers of adenocarcinomas and squamous cancers within the studies examined in this population-level data meta-analysis are unknown and, given the disparity of geographical locations studies, are likely to vary between studies, further contributing to the heterogeneity of results in this subgroup analysis. We acknowledge that these population-level, all-oesophageal cancer data are the weakest and least informative of those specifically included in this systematic review, but feel these data should be included for completeness and to stimulate further research into any associations between statin use and squamous cancer. However, the pooled data on all oesophageal cancers are still consistent with statin use being associated with a lower risk of oesophageal adenocarcinoma.

The main limitation to this meta-analysis is that all included studies are observational (i.e. cohort or case-control). As such, regardless of the sample size involved in this study, it is not completely free of bias and despite the generally high quality of the observational studies included in this systematic review, uncorrected bias may still influence the results. However, the consistency of the results, including lack of heterogeneity, from different geographical locations, health care systems, and research methodologies suggests that unadjusted bias is unlikely to have a major effect on our findings, although it would be important to establish what the association is between statin use and cancer risk in relation to specific other risk factors for OAC such as smoking, obesity, and visceral obesity [8]. In the adenocarcinoma-specific analysis, particularly in the Barrett’s cohorts, the minimal difference between the pooled adjusted and unadjusted ORs, suggests that confounding by indication and residual bias, are unlikely to be contributing to the results.

It was notable that the number and variety of potential confounding variables adjusted varied considerably between the studies and this would seem to be a major focus for future work. Despite the evidence that proton pump inhibitors and NSAIDs [33, 36] are also associated with a reduced incidence of OAC, not all studies provided data to enable accurate adjustment. Laboratory studies and our previous meta-analysis showed that the combination of statins and cyclo-oxygenase inhibitors appeared to be additive in both effects on cell lines [6, 7] and in reducing the incidence of OAC (pooled aOR 0.26 (0.1–0.68)) [12]. However, the more recent studies have not examined or reported the data to enable further verification of this association and it is recommended that any further studies specifically examine this.

Probably the weakest aspect of the current study is the lack of data on statin type (lipophilic or hydrophilic), dose, drug, or duration. Previous meta-analyses [11, 12] were similarly limited in this regard but unfortunately despite the increase in the number of studies and subjects included in the systematic review, the data are still inadequate to draw any firm conclusions. Most studies did not specifically examine this aspect of statin use and when it was reported, the categories used were too heterogeneous for statistical analysis. Where reported there seemed to be a tendency for higher doses and longer durations to be associated with a lower incidence of OAC but any future studies should specifically address these issues.

In conclusion, our updated systematic review and meta-analysis show that statin use is associated with a significantly reduced incidence of oesophageal adenocarcinoma in both Barrett’s cohorts (by 41%) and population-based studies (by 43%). There is also a reduced incidence of all oesophageal cancers in population-based studies (by 18%), although it is possible that this is driven mainly through a reduction in the incidence of OAC rather than OSC; that hypothesis requires further investigation. The inclusion of additional studies has enabled much tighter precision of the estimate of the reduction of OAC associated with statin use in Barrett’s cohorts.

The consistency of the results across multiple studies combined with supportive experimental evidence does suggest that statins may have useful chemopreventative effects in Barrett’s oesophagus. Given the relatively low rate of neoplastic progression in BO (approximately 1 in 300 per year [37, 38]), it may not be cost-effective to merely use statins as chemopreventative agents, but given that the main cause of death in patients with BO is not oesophageal cancer but vascular diseases, it may be prudent to ensure that statins that are definitely prescribed to patients with Barrett’s oesophagus were indicated by vascular risk [39].

Authors’ Contributions TT, YL, and ILPB all jointly conceived the study. TT and ILPB performed the literature review and data extraction. TT and YL performed the data analysis. TT wrote the first draft of the manuscript. ILPB wrote the final version of the manuscript and is the guarantor of the paper. All authors have approved the final manuscript.

No funders were involved in any aspect of this work. The authors have no conflicts of interests.

Compliance with Ethical Standards

Data Availability The datasets analysed during the current study are available from the corresponding author on reasonable request.

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