Wage Difference Between Formal Sector and Informal Sector Jobs; With Special Reference to the Labour Market in Sri Lanka

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Abstract: It is globally understood that wage-based employment structure and wages are a central aspect of the labour force at work. The informal sector is ranging to a broader concept that is difficult to define. The formal–informal wage gap is crucial to understand labour market informality, especially in developing countries with the large informal sectors. The basic model is taken from Mincer (1974), and the study is primarily based on secondary data. The new dummy variable of Job_type and an interactive term were incorporated into the Mincer earning function to analyse wage differences between formal sector and informal sector jobs. The study concludes that there is a wage gap between the formal and informal sector. Moreover, if a person engages in formal sector job with good education qualification and good working experience, he will be entitled for a higher wage rate. Policies that promote education and equal opportunities for workers in both formal and informal sectors would improve earnings for many workers by increasing productivity and incomes.

Keywords: Difference, Formal, Informal, Jobs, Wage.

1. Introduction

It is globally understood that wage-based employment structure and wages are a central aspects of the labour force at work. The informal sector is ranging to a broader concept that is difficult to define. According to ILO, the poor working class and its activities includes the informal sector who are not recognized, recorded, secured, and regulated by public authorities. Maloney (2004) argues that the informal sector cannot be marginalised because it might be a voluntary choice for individuals or businesses due to various factors like aversion of tax, cost of regulations, having greater flexibility and more freedom.

1.1 Problem Statement

Formal employment is recognised of workers with social security benefits provided by the employer, while informal employment refers to its absence. The organized sector consists of considerably regular workers although on an informal basis, workers includes both regular and casual workers. The informal sector include largely casual employment (Papola & Kannan, 2016). The formal–informal wage gap is crucial to understand labour market informality, especially in developing countries with large informal sector. It is widely accepted that both formal and informal sectors play a key role in the global economy.

1.2 The Main Objective of the Study
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The main objective of the study is to investigate the wage difference between the formal and informal sectors and the extent of such difference in Sri Lankan labour market.

2. Literature Review

2.1 Theoretical Review

The wage is measured as returns to education, and it is estimated using two main methods in the literature, that is, the full-discounting method and the Mincerian earnings function. Research references over the years have given preference for researchers to apply the Mincerian method because of its convenience (Psacharopoulos & Patrinos, 2018). Differences in earnings across different socio-economics categories are assessed by adding categorical dummy variables to the basic Mincer earnings. In fact, in many empirical studies, the Mincer earnings function has been modified to incorporate region, union membership status, city size, race, gender, ethnicity, migration status, health status, tenure on the job, and a host of other factors that could affect earnings (Polachek, 2008).

The rate of return to schooling equates the value of lifetime earnings of the individual to the net present value of education costs. The rate of return should have a positive value for an investment to be justified economically and should be higher than the alternative rate of return. An individual point of view, weighing costs and benefits means in investment if the rate of return exceeds the private discount rate (the cost of borrowing and an allowance for risk). The costs incurred by the individual are the foregone earnings while studying, plus any schooling fees or incidental expenses incurred. The private benefits are assessed on how much extra an educated individual earns (after taxes) compared to an individual with less education. Society’s spending on education refers to the social rate of return – for example, money spent on renting buildings and professorial salaries. The estimated rate of return in terms of social attribute refers to the inclusion of the full resource cost of the investment – the direct costs by government and the foregone earnings of students as they invest in their education. The costs are higher in a social rate of return calculation, so comparing relatively to the one from the private point of view, social returns are typically lower than a private rate of return.

Mincer (1974) initiated to derive firstly an empirical equation of earnings over the lifecycle. In Mincer Earning function, at any point (t) in an individual’s lifetime, observed earnings \( Y(t) = wK(t) - s(t)K(t) \) can be depicted as a concave function of one’s labour market experience. Assuming that the schooling phase of investment lasts S years and that on-the-job training declines linearly over the lifecycle, log-earnings are a quadratic function of labour market experience. \( \ln Y(t) = a_0 + a_1S + a_2X + a_3X^2 + \mu \)

- \( a_0 \) is related to initial earnings capacity,
- \( a_1 \) is the rate of return to education (assuming all schooling costs as opportunity-costs)
- \( a_2 \) and \( a_3 \) are related to both the amount and the financial return to on-the-job training

The Mincerian earnings function has been the subject of controversy in the literature (Psacharopoulos and Layard, 1979; Heckman, Lochner, and Todd, 2006). One issue with the Mincerian method of estimating returns to education is missing variables, e.g. ability bias. Griliches (1977) analysed the issue many years ago. He found that the bias is small or negative. Adding more variables to the equation will not solve the problem and might add other biases (Patrinos, 2016).

The earnings premium associated with the level of education suggests that productivity increases as people acquire additional qualifications. An alternative view is that earnings increase with education due to credential effects. This refers to the idea that higher levels of schooling are associated with higher earnings, not because they directly raise productivity, but because they certify that the worker is likely to be productive. In this sense, education merely sorts workers according to their unobserved attributes; it does not necessarily augment their intrinsic productivity. Moreover, The Mincerian method gives private returns, whereas the full discounting method can give private and social returns.
2.1 Empirical Review

Economists have studied the wage gap between formal-informal workers to establish the facts whether informal workers earn less than their formal counterparts. Alternative theories have been put forward to explain this phenomenon. The traditional labour theory argues that workers choose the informal sector because they are limited to the entry of formal sector due barriers (Fields, 1990; Harries & Todaro, 1970).

Henley, Arabsheibani, and Carneiro (2006) found from the study contents in Brazil that informal workers on average tend to earn less than their formal counterparts. Tannuri and Pianto (2002) point out in the entire wage distribution in Brazil that a positive wage gap is in favour of formal workers. The study results done in Brazil, South Africa and Mexico by Bargain and Kwenda (2009) also indicate a significant formal–informal wage gap at lower quantiles. Sookram and Watson (2008) say that formal sector workers earn more than informal sector workers and work experience is the important determinant of the wage gap in Caribbean islands Trinidad and Tobago.

Polachek (2008) confirm the positive correlation between schooling and earnings based on the data of 2003 Current Population Survey to assess rates of return to education across occupational categories in the U.S. labour market. They find that “additional schooling” has a positive impact on the weekly earnings of men and women in both white- and blue-collar occupations.

Gong and van Soest (2002) apply quarterly panel data from Mexico in the analysis of the wage differentials between formal and informal sectors. They use random effects wage regression to explain the wage formation and differentials, thus managing likely selection bias because of unobserved time-invariant heterogeneity which affects both wages and sector choice. This is the first such study to determine simultaneously wages and sector choice in one dynamic panel data setting. Using Heckman (1981)’s Monte Carlo simulated maximum methodology, Gong and van Soest point out that age significantly affects formal sector wage, not the informal sector wage. Although much higher returns to education show in the formal sector, returns seem positive in both sectors. The lagged labour market state has no effect on wages and random effects are insignificant in the wage determination process. For testing formal-informal wage differentials divide in Argentina, Pratap, and Quintin (2006) prefer to propensity matching score methodology (PSM) to find the sample selection problem often inherent in such analyses. As in line with other studies, they find a 25 percent formal wage premium using standard OLS estimation, controlling for individual and establishment characteristics. However, once semi-parametric methods are matched in observation of similar workers, Pratap and Quintin find no evidence of a formal-sector wage premium and thereby reject the segmented formal-informal Argentinian style labour markets theory. They evaluate the firm importance of controlling firm size in their last section of analysis, worker characteristics which are unobserved may affect both selection decision and wages and the value of other pecuniary and non-pecuniary benefits of a job.

The informal-formal wage gap is examined by Bargain and Kwenda (2009) using large panels in Brazil, Mexico, and South Africa. The study is twofold. First, usual measures of wage are adjusted for the taxes paid in the formal sector likely to cause overestimation of the same sector wage premium. Secondly, time-invariant unobserved heterogeneity is accounted for by using fixed-results quantile regression estimation of Koenker (2004) and Canay (2011). The sample is included only urban male aged 15-65 not having in any form of education, working as unpaid family or public worker. Females are excluded from the sample given that most are engaged in unpaid family work and accounting for selection into labour market is not yet standard in quantile regressions. The findings reveal a similar distributional pattern of informal wage penalty across all countries, as resulting informal wage gap prevails mostly in lower earnings quantiles and disappears at the top quantiles. Blunch (2011) adds to the existing facts by examining the magnitude and determinants of formal-informal sectors earnings gap in Serbia even amidst in the recent International Financial Crisis. The empirical analysis is carried out and referred to across four alternative measures of informality (firm registration, labour contract, benefit
receipts and firm size) and two time periods of 2008 and 2009. In particular, Blunch first estimates the raw formal-informal sectors earnings gap through Mincer wage regressions using ordinary least squares, then applies overall and detailed Blinder-Oaxaca decompositions to the observed earning gaps. The findings evince a large formal-informal sectors earnings gap which somehow appears to decrease following the onset of the crisis. However, the gap does not show a noticeable change when controlled for observable characteristics.

### 3. Research Methodology

The basic model is taken from Mincer (1974). The study is primarily based on secondary data and the sample size is equal to 1705.

\[
\ln Y = \alpha_0 + \alpha_1 S + \alpha_2 X + \alpha_3 X^2 + \mu
\]

where \( \ln Y \) is the log value of the rate of return to education, \( S \) is the years of schooling, \( X \) is equal to total years of experience and \( X^2 \) is the square value of experience.

In this study new dummy variable, \( \text{Job\_type} \) (formal sector or informal sector jobs) and an interactive term were incorporated into the Mincer earning function to analyse wage difference between formal sector and informal sector jobs.

\[
\ln Y = \alpha_0 + \alpha_1 S + \alpha_2 X + \alpha_3 X^2 + \alpha_4 D + \alpha_5 (SD) + \mu
\]

\( D \) is a dummy variable (formal sector and informal sector) and \( SD \) is an interactive term (years of schooling multiply by dummy variable). \( \alpha_4 \) measures the effect on intercept (Average) and \( \alpha_5 \) measures effect on rate of returns to education (wage) When \( D = 0 \) (person from informal sector)

\[
\ln Y = \alpha_0 + \alpha_1 S + \alpha_2 X + \alpha_3 X^2
\]

When \( D = 1 \) (person from formal sector)

\[
\ln Y = (\alpha_0 + \alpha_4) + (\alpha_1 + \alpha_5) S + \alpha_2 X + \alpha_3 X^2
\]

The following variables were selected from the data set given to estimate the model.

- **EPF** - Employer has registered under EPF or Inland Revenue Department
- **Age** - Age of the person
- **Formal** - Maintain accounts
- **Income** - Total income from all sources in both primary and secondary jobs
- **Education** - Years of schooling

| Variable | Label | Type of the variable | Derivation of variable |
|----------|-------|-----------------------|------------------------|
| \( \ln Y \) | Log\_Income | Continuous | Log value of the income |
| \( S \) | Education | Continuous | Years of schooling |
| \( X \) | Experience | Continuous | Age - 5 - 6 |
| \( X^2 \) | Square value of the variable experience | Continuous | \( X^2 \) |
| \( D \) | Type of the job | Dummy | \( D = 1 \) if EPF = 1 and if formal = 1 \( D = 0 \) if EPF = 0 and if formal > 1 |
| \( SD \) | Interactive term | Continuous | \( S \cdot D \) |

### 3.1 The Rationale for Constructing the Dummy Variable (\( D = \) Type of the Job)

As per the definitions given by the ILO and Census and Statistics Department, the formal sector jobs refer to any employer who maintains formal accounts, contributing to EPF, paying taxes and provide many social security benefits to the employees. In contrast, informal sector jobs refer to people who work on casual basis and they do not maintain any formal accounts, paying
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taxes or contributing to EPF. Thus, by considering the given data set, the variable “D” is constructed based on both formal and EPF variables.

3.2 Data Cleaning – Handling Missing Values

Improper handling of missing values will lead to draw an inaccurate inference about the data. In fact, the researcher may have two options either to drop the missing data or replace them by mean or mode value. In this study, the number of missing values of the constructed dummy variable (D) is 1295 and it is quite a large proportion out of the total sample size. As mentioned above, there are two possible remedies either to drop the missing values or replace them by the mode value (because D is a dummy variable). Since, the main objective of the study is to detect whether is difference in wage between formal and informal sector jobs, replacing 1295 missing values will lead to draw a biased result. Even after deducting the missing values, the total sample is equal to 1705 which is sufficient to run log-linear regression analysis. Thus, the best possible remedy is to omit missing values from the data set.

Regression Analysis

Regression model: \( \ln Y = \alpha_0 + \alpha_1 S + \alpha_2 X + \alpha_3 X^2 + \alpha_4 D + \alpha_5 (SD) + \mu \)

4. Data Analysis and Interpretation

4.1. Descriptive Statistics

Table 4.1: Descriptive Statistics of variables.

| Variable | Mean | Std. dev | Std. error | [95% Conf. Interval] | Min | Max |
|----------|------|----------|------------|----------------------|-----|-----|
| S        | 10.162 | 3.432 | 0.082 | 9.388 | 9.709 | 0 | 17 |
| X        | 23.352 | 14.099 | 0.390 | 22.256 | 23.669 | -1 | 66 |
| X2       | 744.030 | 769.879 | 19.646 | 706.775 | 786.839 | 0 | 4356 |
| ln_Y     | 10.084 | 0.699 | 0.017 | 9.956 | 10.022 | 6.908 | 13.122 |
| D        | 0.581 | 0.494 | 0.012 | 0.858 | 0.604 | 0 | 1 |
| SD       | 6.188 | 5.754 | 0.140 | 5.914 | 6.461 | 0 | 17 |

Source: Author constructed, 2020.

The mean value of the log income is 10.08 and its standard deviation is 0.699. The mean of an interactive term (SD) is 6.188 and its standard deviation is 5.754. Education and experience report 10.162 and 23.352 mean values, respectively.

Chart 4.1: Mean comparison of wage between formal sector vs informal sector.

Source: Author constructed, 2020
4.2 Interpreting the Results of an Uncorrected Model

R2 and Adj. R2 indicate the Coefficient of Determination and its adjusted value, respectively. These explain the proportion of total variation of Y explained by the regression, but Adj. R2 is more accurate than R2. As per the regression results R2 = 0.2510 and Adjusted R2 = 0.2492 which means that the independent variables, explain 25% of the variability of the dependent variable (ln_Y) which is quite insufficient. However, the regression model is statistically significant, F (5, 1699) = 114.12, p = 0.0000. This indicates that, overall, the model applied can statistically significantly predict the dependent variable, wage.

\[ \ln Y = 9.3167 + 0.0086(S) + 0.0348(X) - 0.0006(X^2) - 0.6587(D) + 0.1012(SD) \]

Since \( a_1 \) is not statistically significant (p=0.271>0.05), the coefficient of S \( (a_1 + a_5) \) is calculated only by considering the \( a_5 \).

\[ \ln Y = 8.658 + 0.1098(S) + 0.0348(X) - 0.0006(X^2) \]

The significant positive coefficient for SD \( (p = 0.000 < 0.05) \) indicates that there is a significant difference of rate return to education (wage) between formal and informal sector jobs. A person engaged in formal sector jobs receives 10% (ln Y increase by 0.1012 when D=1) higher wage.

Source: Author constructed, 2020.
rate than the person who is engaged in informal sector job. The magnitude of the effect over the rate return to education is much higher for person engaged in formal sector job. If a person engages in informal sector job; education does not affect for determining wage rate, but increase in years of experience will increase the wage rate by 3%. But, if a person engages in formal sector job; increase years of schooling (S) by one unit will increase wage rate by 10% (0.1012) and increase in years of experience will increase the wage rate by 4% (0.0348). Thus, if a person engages in formal sector job with good education qualification and good working experience, he will be entitled for higher wage rate. The results of the study indicate that the education and work experience will increase magnitude over wage rate for formal sector job holder. Moreover, work experience has positive impact over wage rate of informal sector job holder, whereas education has no any significant impact.

4.3. Second Order Tests

Table 4.3: Test results of Multicollinearity and Heteroscedasticity.

| Variable | VIF   | 1/VIF  |
|----------|-------|--------|
| SD       | 14.1  | 0.070928 |
| X2       | 12.53 | 0.079837 |
| X        | 12.35 | 0.080995 |
| D        | 9.94  | 0.100594 |
| S        | 3.26  | 0.306289 |
| Chl2 (1) | 11.24 |        |
| Prob > chl2 | 0.0008 |    |

Source: Author constructed, 2020.

Since VIF value of SD is 14.10 and X2 is 12.53; Multicollinearity between X and X2 and between S and SD are at harmful level. Prob > CHI2 = 0.0008 indicates that the null hypothesis is rejected. This means the model estimated above is contaminated with Heteroscedasticity.

Adjustment for the problem of Heteroscedasticity; White correction of Error Variance.

The robust was conducted to adjust the problem of Heteroscedasticity and coefficients are slightly differ than the uncorrected model.

4.4 Mathematical Adjustment for the Problem of Multicollinearity

\[
\ln Y = a_0 + a_1S + a_2X + a_3X^2 + a_4D + a_5(SD) \\
X^2 = \delta_0 + \delta_1X \quad (\text{reg } X^2 X) \\
SD = \delta_2 + \delta_3D \quad (\text{reg } SD D) \\
\ln Y = \theta_0 + \theta_1S + \theta_2X + \theta_3e^2 + \theta_4D + \theta_5e^3
\]

\[
a_0 = \theta_0 - a_3\delta_0 - a_5\delta_2 \\
a_0 = 9.5894 - (-0.0006 \times (-451.2403)) - (0.1012 \times (-1.23e-13)) = 9.5894 - 0.2707 - 0 = 9.3187
\]
Regression results of corrected regression model;

\[
\ln Y = 9.3187 + 0.0086S + 0.0345X - 0.0006X^2 - 0.6587D + 0.1012SD
\]

Since \( \alpha_1 \) is not statistically significant (\( p=0.271>0.05 \)), it is omitted from the estimated model.

\[
\ln Y = 9.3187 + 0.0345X - 0.0006X^2 - 0.6587D + 0.1012SD
\]

### Table 4.4: Comparison of Corrected Model with uncorrected model.

|       | Estimates of uncorrected Model | Estimates of corrected Model |
|-------|--------------------------------|-----------------------------|
| \( \alpha_0 \) | 9.3167                          | 9.3187                      |
| \( \alpha_1 \) | 0.0085                          | 0.0086                      |
| \( \alpha_2 \) | 0.0348                          | 0.0345                      |
| \( \alpha_3 \) | - 0.0006                        | - 0.0006                    |
| \( \alpha_4 \) | - 0.6587                        | - 0.6587                    |
| \( \alpha_5 \) | 0.1012                          | 1.1012                      |

**Source:** Author Constructed, 2020.

### 4.5 Interpreting the Results of a Corrected Model

When \( D = 0 \) (person from informal sector)

\[
\ln Y = \alpha_0 + \alpha_1S + \alpha_2X + \alpha_3X^2 \ln Y = 9.3187 + 0.0086S + 0.0345X - 0.0006X^2
\]

Since \( \alpha_1 \) is not statistically significant (\( p=0.271>0.05 \))

\[
\ln Y = 9.3187 + 0.0345X - 0.0006X^2
\]

When \( D = 1 \) (person from formal sector)

\[
\ln Y = (\alpha_0 + \alpha_4 ) + (\alpha_1 + \alpha_5)S + \alpha_2X + \alpha_3X^2 \ln Y = 8.660 + 0.1098(S) + 0.0345(X) - 0.0006 (X^2)
\]

Since \( \alpha_1 \) is not statistically significant (\( p=0.271>0.05 \)), the coefficient of \( S (\alpha_1 + \alpha_5) \) is calculated only by considering the \( \alpha_5 \).

\[
\ln Y = 8.660 + 0.1012(S) + 0.0345(X) - 0.0006 (X^2)
\]

The significant positive coefficient for \( SD (p = 0.000 < 0.05) \) indicates that there is a significant difference of rate return to education (wage) between formal and informal sector jobs. A person engaged in formal sector jobs receives 10\% (\( \ln Y \) increase by 0.1012 when \( D=1 \)) higher wage rate than the person who is engaged in informal sector job. The magnitude of the effect over the rate return to education is much higher for person engaged in formal sector job.

If a person engages in informal sector job; education does not affect for determining wage rate, but increase in years of experience will increase the wage rate by 3\%. But, if a person engage in formal sector job; increase years of schooling \((S)\) by one unit will increase wage rate by 10\% \((0.1012)\) and increase in years of experience will increase the wage rate by 3\% \((0.0345)\). Thus,
if a person engages in formal sector job with good education qualification and good working experience, he will be entitled for higher wage rate. The results of the study indicate that the education and work experience will increase magnitude over wage rate for formal sector job holder. Moreover, work experience has positive impact over wage rate of informal sector job holder, whereas education has no any significant impact.

**Log-Linear Regression Model** – an extension to a mincer earning function is used to analyse whether the return on education (wage) is differed between formal and informal sector jobs.

\[
\ln Y = a_0 + a_1S + a_2X + a_3X^2 + a_4D + a_5(SD)
\]

\[H_0 \; ; \; a_4 = a_5 = 0\]

Since, sig value for \(a_4\) is 0.000 (p = 0.000 < 0.05) and sig value for \(a_5\) is 0.000 (p = 0.000 < 0.05), \(H_0\) is rejected, concluding that there is significance difference of rate of return to education between formal and informal sector jobs. A person engaged in formal sector jobs receives 10% (ln Y increase by 0.1012 when D =1) higher wage rate than the person who is engaged in informal sector job.

**4.6 Estimates Of Corrected Regression Model**

Since \(a_1\) is not statistically significant (p=0.271>0.05)

\[
\ln Y = 9.3187 + 0.0086S + 0.0345X - 0.0006X^2 - 0.6587D + 0.1012SD
\]

When \(D = 0\) (person from informal sector)

\[
\ln Y = 9.3187 + 0.0086S + 0.0345X - 0.0006X^2
\]

When \(D = 1\) (person from formal sector)

\[
\ln Y = (a_0 + a_4) + (a_1 + a_5)S + a_2X + a_3X^2 \ln Y = 8.660 + 0.1098(S) + 0.0345(X) - 0.0006 (X^2)
\]

The study estimates that there is significant difference of rate return to education (wage) between formal and informal sector jobs. A person engaged in formal sector jobs receives higher wage rate than the person who is engaged in informal sector job. If a person engages in informal sector job; work experience has positive impact on wage rate. Moreover, if a person engages in formal sector job with good education qualification and good working experience, he will be entitled for higher wage rate. Thus, the magnitude of the effect over the rate return to education is much higher for person engaged in formal sector job.

**5. Conclusion and Recommendations**

The study concludes that there is a wage gap between the formal and the informal sector. The findings of this paper suggest that human capital factors i.e., education and experience have a key role in improving the wages. The higher educated and more experienced workers are likely to have higher earnings in the formal sector. Policies that promote education and equal opportunities for workers in both formal and informal sectors would improve earnings for many workers by increasing productivity and incomes.

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