Living in the tide of change: explaining Japanese subjective health from the socio-demographic change

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INTRODUCTION

Studies of culture change have paid attention to the change in the level of national values (Hofstede, 2001; Allen et al., 2007; Bilsky et al., 2011; Hamamura, 2012; Twenge et al., 2012), but few studies have been conducted on the relationship between the change and health of residents (Oishi et al., 2011). We argue that change in the socio-demographic conditions would undermine individuals’ psychological health, because whereas human well-being is a function of habituated behavioral tendencies fitted best to accustomed context, change will bring alterations to the requirements of the external environment. We believe such process is apparent across national boarders. In this study, we focus on the typical case of socio-demographic change which took place in Japan.

After the economic crisis known as the collapse of the bubble economy,1 Japan has witnessed increasing societal problems, and this has led to investigations on the relationship between culture change and health (Norasakkunkit et al., 2012). In this study, we take the perspective that Japanese culture, especially its socio-demographic context is becoming more non-traditional (Hamamura, 2012), and the change in context from 1990 to 2010—the two decades of extremely low rate of new employment opportunities, spread of new liberalism, and market globalization (Zielenziger, 2006)—would negatively explain the health of Japanese adults.

THE WORLDWIDE TIDE OF CHANGE

The world shift toward Gesellshaft (Tönnies, 1887/1988) or socio-demographic condition common in individualistic societies (Greenfield, 2013) is one of the largest contextual changes surrounding recent human ecology. Starting from economically developed countries, urban population, internet users, one person households, divorce, and national affluence are increasing throughout the world. Conceptually, we regard these changes as indicating the change toward increased Gesellshaft socio-demographic condition. Figure 1 shows a world summary of the changes, summarized from the database of OECD (Organization for Economic Cooperation and Development: http://www.oecd-ilibrary.org), United Nations Economic Commission for Europe (http://w3.unece.org/pxweb/), and World Bank (http://data.worldbank.org/indicator), from 1990 to 20102 (OECD, 2014; United Nations Economic Commission for Europe, 2014; World Bank, 2014). Across the nine indicators, average ecological correlation with time (Hamamura, 2012) was $r = 0.94$ ($p < 0.05$), suggesting that these variables are changing worldwide with time.

1A surge in the market caused by speculation regarding a commodity which results in an explosion of activity in that market causing vastly over inflated prices. The prices are not sustainable, and a bubble is usually followed by a crash in the affected sector. In the Japanese case, land prices first inflated and then crashed.

2We selected nine variables based on Greenfield (2013) and Hofstede (2001), as indicators of Gesellshaft or socio-demographic conditions (Greenfield, 2013) related to individualism (Vandello and Cohen, 1999; Yamawaki, 2012) of the country.
The change stimulated national development in many countries, fulfilling our needs of the modern life style (Inglehart and Baker, 2000). However, during the same period, many developed countries have also been caught up in waves of recession, leading to disparity in the population. During this period, worldwide economic stagnation occurred in East Asian countries in early 1990s, and in European countries during the late 2000s. There were some serious health consequences for marginalized group, such as unemployed, uneducated, or poor individuals (Stuckler et al., 2009) during the time. However, economic turmoil may also affect the health of population at large, via the change in peoples’ socio-demographic conditions.

**PSYCHOLOGICAL CONSEQUENCES OF CHANGE**

Historically, the socio-demographic change toward Gesellshaft secured free market and wealth leading to material possession and happiness (Diener, 2000; Inglehart et al., 2008). However, since socio-demographic change, by definition, takes place first in the external environment, it always precedes psychological change (Bardi and Goodwin, 2011). Because human cultural adaptation is built even at the level of automatic behaviors learned in accustomed daily activities (Kitayama and Park, 2010), change in the socio-demographic context would impose an implicit challenge to the efficiency of acquired pattern of behaviors. If acquired behaviors were rendered useless to pursue valued goals in life (Oishi et al., 1999b), our well-being would be compromised. In other words, our health is usually sustained by the culture we live in (Diener, 2000), Draguns and Tanaka-Matsumi (2003), and Triandis (2007), but for this very reason, socio-demographic change undermining the original can cause mental health problems in the population (Kashima, 2013), by creating unaccustomed mandates (Kitayama et al., 2009).

If the transition had been carried out smoothly, involving well-structured institutions and formal education that supports people to catch up with the new requirements, the change into Gesellshaft socio-economic condition may foster independence and also preserve healthy social relations (Kagitcibasi, 2005). However, when the transition is rapid, imminent, allowing no choice but to engage in global competition to survive, general population would fail to adapt. Further, if the original culture was encouraging interdependent ways of life, rapid change into Gesellshaft socio-demographic condition might require extra mental effort for the residents (Ogihara and Uchida, 2014).

Possible disadvantages due to such change may range from psychological to social aspects of their health. Specifically, their
general health might be compromised because their working conditions become worse in response to economic stagnation. During this period, full-time employees were required to oblige by meeting an intensified workload, and they faced a doubled unemployment rate (Statistics Bureau, 2014). Their life satisfaction might be lowered because subjective well-being can be compromised when valued life domains fail to satisfy individuals due to economic turmoil (Oishi et al., 1999a; Diener, 2000; OECD, 2013). If rapid change requires the members to be alone and exercise new self-ways to be accepted, change might compromise one’s self-worth, or the fundamental motivation to belong (Baumeister and Leary, 1995) and one’s achievement/acceptance of socially desirable self (Tafarodi and Swann, 1995). Perceived social support might also be compromised as change may increase one’s need to relocate and be mobile in the new context (Oishi et al., 2007). Isolated from supportive others, one’s perceived social support may decline.

In East Asian countries, the outcome of the change has been observed as a decrease in average psychological health. For example in Japan, people have traditionally regarded interpersonal harmony as their central meaning of happiness (Uchida and Kitayama, 2009). Now they have to leave their close families, elder parent or friends in order to attain new position or decent job, sometimes they are involved in an aggressive competition or candid quarrel against their colleagues in order to compete for status and limited wage. Otherwise, they may give up making new family out of inflated partner choice and economic/business reasons. These events might increase the likelihood of interpersonal concerns or disharmony. Especially for those having interdependent self, the change might undermine their well-being, or severely limit the fulfillment of the motivation to belong.

Given these disadvantages stemming from the glitch at socio-demographic level, some people may survive the influence using specific forms of positive strength. As one of such strengths, we focused on interdependent happiness3 (Hitokoto and Uchida, 2014). Interdependent happiness is a collectively shared concept of happiness (Uchida et al., 2014) among the members of interdependent cultures. According to the previous studies (Uchida and Kitayama, 2009; Uchida and Ogihara, 2012), Japanese, compared to European American counterparts, share the meaning of happiness as relational. Whereas this type of happiness is held central to the subjective concept of happiness among certain cultural members, the conception of happiness as relational is considered common across all human being. Because interdependent happiness involves relational harmony, quiescence and ordinaryness (Hitokoto and Uchida, 2014), these positive meanings would be more central to residents’ well-being in the Gemeinschaft societal condition. If the change is an avalanche toward Gesellschaft socio-demographic condition, then the traditional interdependent happiness would be a positive buffer against the trend. Therefore, among the traditionally Gemeinschaft societal members, the strength of preserving harmony among close others may find the negative influence of change. Particularly, in Japanese case, interdependent happiness will predict health positively while the change will not.

**AFTERMATH OF RECESSION IN JAPAN**

During the 1990’s, Japanese companies faced a collapse of the collectivistic working style, together with the collapse of the real estate *bubble economy*. In the workplace, companies abandoned the lifetime employment and seniority system, in exchange for thrust management1. The push for globalization economy: transparency, international competition, and global communication increased. Even after 2000, the aftermath of this decade and further economic depression delayed Japan’s economic recovery. Even more tragic was the fact that, although such change was apparent, formal education continued to encourage *generalists*: all-rounders capable of managing multiple roles in a single company, thus best functional under the long-lasting employment system (Toivonen et al., 2011)2, and the major companies preferred naïve freshmen over slightly older professionals as their main labor force.

Coinciding with such tension was an increase in health problems. The decades after 1990s are marked as a time when Japan showed unprecedented deterioration in mental health statistics (Kawakami et al., 2006). Following the *bubble* collapse, over 30,000 workers per year committed suicide. Specifically, the suicide rate of working men in their 30’s and 50’s increased. The number of outpatients diagnosed with mood disorders doubled between 1990 and 2010. The number of working adults suffering from depression increased, with more people suffering from depression than cancer or diabetes. From the late 1990’s to 2010, applications for insurance compensation to workers were dominated by requests from patients with mental disorders (Japanese Association of Schools of Certified Psychiatric Social workers, 2012), and industrial safety and health issues became focused on regulating overwork (Health, labour and wage statistics associations, 2012). Given this situation, the Ministry of Health, Labor, and Welfare listed mental disorders as one of the big five diseases among Japanese adults in 20116. The decades after 1990 were a time when the lack of fit between traditional interdependence and the requirements of the new liberal, global economic, individualistic environment became apparent, with critical consequences for Japanese health. Kitanaka (2012) vividly portrays these adverse psychological profiles of Japanese workers and a “society in distress” through narrative data.

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3Hitokoto and Uchida (2014) developed the interdependent happiness scale (IHS), which consists of items pertaining to relational harmony, quiescence, and ordinariness, distinguishable from other types of happiness.

4In households, single living and divorce increased, and the average number of family members decreased rapidly during this period (Statistics Bureau, 2010).

5Such a mysterious mixture of change and tradition can be perceived at a glance in a single workplace where elders endorse a traditional value orientation in contrast to the young (Kawakami and Glazer, 2010), or in the fact that elders are more collectivistic (Matsumoto et al., 1996).

6In addition to these mental health issues, other societal problems also increased, such as social withdrawal, lowering of the birth rate, substance addiction, pathological gambling, school bullying, and teachers’ mental issues, juvenile suicide, unemployment, economic disparity and the “lonely death (Ko-doka-shit)” of elders (Japanese Association of Schools of Certified Psychiatric Social workers, 2012).
HYPOTHESIS
What are the psychological consequences of economic stagnation accompanying demographic changes away from the traditional way of life? In this study, we used a national case approach and focused on Japan as a country where people have traditionally lived in *Gemeinschaft* socio-demographic condition, but now caught up in a rapid change as a consequence of the 1990s (Toivonen et al., 2011).

Specifically, we hypothesized that among those working adults who underwent large change toward individualistic socio-demographic condition, their psychological health would be more compromised than those who underwent small change (Figure 2). We tested this hypothesis by comparing different regions that systematically differ in the amount of change.

**Dependent variables**
We measured adults’ health according to the indicators of (1) general health, (2) life satisfaction, (3) self-esteem, and (4) perceived social support. We hypothesized that the change would negatively predict general health (Hypothesis 1). Also, change will negatively predict life satisfaction (Hypothesis 2), self-esteem (Hypothesis 3), and perceived social support (Hypothesis 4). We divided social support into support provided from one’s community and support provided from out of one’s community, in order to explore the possible difference between these two. Finally, interdependent happiness would positively predict psychological health (Hypothesis 5).

**Independent variables**
In this study, we use the term “change” to refer to a set of changes in socio-demographic variables related to societal level individualism-collectivism (Vandello and Cohen, 1999), such as family size or divorce. Specifically, the change is captured as (1) a cross-temporal change during 1990–2010, and (2) a change which took place at the level of socio-demographic condition. Therefore, we operationalized change as the difference in scores between two or more fixed points in time (Hofstede, 2001; Zielenziger, 2006; Allen et al., 2007; Inglehart et al., 2008). We examined change using ecological correlations between our indicators discussed below and time (Hamamura, 2012).

![FIGURE 2 | Hypothesized multi-level model predicting aspects of health from prefecture, city, and individual level predictors. Arrows indicate positive explanatory weights, while broken arrows indicate negative explanatory weights.](image-url)
We measured socio-demographic conditions both at the meso (i.e., prefecture) and micro (i.e., city) levels (Iwata, 2006), and used change at those levels as additional contextual explanatory variable.

In order to compare groups of Japanese adults varying in their degree of change, we used prefecture difference (Plaut et al., 2002; Kitayama et al., 2006). Yamawaki (2012) measured at Japanese prefecture level: the Japanese Collectivism Scale (JCS). The JCS is a sum of standardized scores of (1) divorce to marriage ratio (reversed), (2) percentage of households with three generations living together, (3) percentage of elderly people (aged 65 + years) living alone (reversed), (4) percentage of nuclear family households (reversed), and (5) percentage of people living alone (reversed) in 2006.

Communities also vary by how people make living within their ecologies (Uskul et al., 2008) and they show sizable group differences (Gordon, 2003; Uskul et al., 2008; Fukushima et al., 2011). We considered worthwhile to examine the impact of change at smaller levels than prefectures from the viewpoint of community-based policy making (Raudenbush, 2003; Uchida and Takemura, 2012). We created a JCS city level variable (JCSCL)8 by aggregating the JCS indicators at the city level.9

8Prefectures in Japan are the officially defined local governments that support citizens’ efficient political administration and development of an area. They originate from the boundaries between local warlords during the Edo era that lasted for 300 years, until late 19th century. They are treated as inclusive geographic regions that have distinctive natural and unique historical aspects such as resources, main industries, or dialects.

9Based on the argument by Na et al. (2010), and Hofstede et al. (2010), we expected systematic differences in the reliabilities of JCS and JCSCL. Using various behavioral measures that replicated Japanese and European American differences, Na et al. (2010) found these measures to be uncorrelated with each other within nation, but correlated at the national level, demonstrating a cross-level difference in the reliability. Hofstede et al. (2010) also used his value survey module in a survey of Brazilian states, and poorly replicated his cross-national level factor structure at the cross-state level. If these studies are showing that national level indicators do not always hold at lower levels of analysis, collective behaviors such as percentage of people living with parents or divorce rate may be strongly correlated with each other, thus reliable, at most at the national level (i.e., individualism score and living with parents \( r = -0.67 \); and divorce rate \( r = 0.76 \), Hofstede, 2001). But, once the same societal behaviors are collected at the lower prefecture or city level, the same indicators may lose their reliability accordingly because of the loss in “appropriate” level of analysis. Therefore, if we calculate JCSCL at the city level, we should expect even lower reliability and validity than Yamawaki (2012)’s JCS. Based on this reasoning, we tested the reliability and validity of JCS and JCSCL, hypothesizing that the former would be larger than the latter.

Cities in Japan are officially the most basic individual unit of local government, and their range is defined by population size and concentration of residences. Normally a city is led by a mayor, and each unit has its own board of education to administer formal schooling. While prefecture is rooted in a historical and geographical divide, the city is one of the closest and most rooted living units for residents in Japan. Traditionally, neighbors sharing the same city had specific as well as unwritten norms to oblige and help each other on ceremonial occasions such as marriages and funerals, in exchange for a daily commitment to district matters through assigned roles for each house. Even today, households often have some assigned roles in city affairs such as gatherings or athletic meetings, and people interact face-to-face with their neighbors and most children are formally educated at local schools located within each city.

We evaluated societal level explanatory effects after controlling for basic individual difference variables. For those, we controlled for personality traits (Cloninger, 1987), gender (Kashima et al., 1995), age (Matsumoto et al., 1996) and income (Snibbe and Markus, 2005) in the analysis. As for income, considering the gender distribution of labor altering the income opportunity of adults in Japan, we measured own income independent of partner’s income, and controlled these two in the analysis.

In order to test our model, we first calculated the extent to which each prefecture and each city have changed using archival data (Study 1). We then used those change scores to explain the health of residents living in diverse areas in Japan using social survey (Study 2).

**STUDY 1: MEASUREMENT OF PREFECTURE AND CITY LEVEL CHANGE**

First, we examined how JCS (Yamawaki, 2012) changed between 1990 and 2010. Second, JCSCL and their changes were measured. We tested reliability and validity using criteria that are both available at the city level. Because the change after 1990s should have affected both urban and rural regions during this special period of economic disaster and globalization (Norasakkunkit et al., 2012), we expected change in both urban and rural prefectures.

**METHOD**

**Prefecture change**

We collected archival data on five indicators of the JCS (Yamawaki, 2012) for 47 prefectures in Japan from Japan’s official census statistics site (http://www.e-stat.go.jp/SG1/estat/eStatTopPortaleE.go) at 1990, 1995, 2000, 2005, and 2010. We used JCS at 2010 as a prefecture level Individualistic-Collectivistic socio-demographic condition score in this study10. The five indicators included divorce to marriage ratio, percentage of households with three generations living together, percentage of elderly people (aged 65 + years) living alone, percentage of nuclear family households, and percentage of people living alone (reversed). A total of 1,175 (5 indicators × 47 prefectures × 5 time points) data points were collected. We calculated internal consistencies for every sampled year. We aggregated the standardized subtraction scores of each indicator between 2010 and 1990 as an indicator of change.

**City change**

Five indicators of JCS were collected at the city level (\( n = 1750 \)) from the same archive11. We tested reliability of the JCSCL by calculating internal consistency for every sampled year. We evaluated validity of the JCSCL using correlations with the city level

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10JCS calculated at 2010 strongly correlated with JCS at 1990 (\( r = 0.96, p < 0.001 \)), as well as JCS at 1980 (\( r = 0.98, p < 0.001 \)). The results would indicate, that even though absolute value has changed, the relative position of prefectures on JCS was similar way before the target time period.

11Because “percentage of three generation households” was not available at this level from the same archive, we excluded it from 1990 and 1995, allowing 4 items instead of 5, in constructing JCSCL in the 1990’s. This item reduction resulted in systematic lowering of the reliability of the JCSCL during the 1990’s.
variables of (1) taxable income, (2) percentage of adults working in primary industry, (3) percentage of adults working in tertiary industry, (4) move-in rate, and (5) move-out rate, following past research (Hofstede, 2001; Yamawaki, 2012). Then, we calculated cross-temporal change of JCSCL.

RESULTS AND DISCUSSION

Reliability of the JCS (prefecture level)
The reliability (internal consistency) of the JCS indicators in each sampled year was marginal to acceptable ($0.68 < \alpha < 0.84$) as in the original study (Yamawaki, 2012). The JCS and its change score correlated positively ($r = 0.34, p < 0.05$), indicating prefectures with collectivistic socio-demographic condition showed more change during this period.

In Figure 4, a bar describes the sum of standardized change across indicators for one prefecture, or the degree to which a prefecture changed into individualistic socio-demographic condition in comparison to others. The rate of change differed across prefectures.

Reliability and validity of the JCSCL (city level)
As in Yamawaki (2012), the JCSCL also showed geographic similarity between the neighbors; for example, those cities located around the Sea of Japan coastal area are generally collectivistic, pointing to the possibility that city level socio-demographic condition is clustered geographically (Figure 5). The reliability of the JCSCL was lower ($0.39 < \alpha < 0.62$) than that of the JCS in every sampled year. Similar to the JCS, there was also a declining trend of the reliability over time. The validity coefficients were $r = -0.11 (p < 0.001)$ with income, $r = 0.11 (p < 0.001)$ with workers in primary industry, $r = -0.44 (p < 0.001)$ with workers in tertiary industry, $r = -0.36 (p < 0.001)$ with move-in rate, and $r = -0.40 (p < 0.001)$ with move-out rate, all in expected directions. These indicate reasonable relationships between criteria variables and the JCSCL; yet, they show lower validity coefficients than those of the JCS (Yamawaki, 2012).

Changes of city socio-demographic condition
The second column of Table 1 shows the correlation between socio-demographic indicators and time at city level. All indicators changed in the direction of increased individualistic socio-demographic condition, with percentage of nuclear family showing weaker correlation than other indicators. The changes in the JCSCL scores also varied between communities. The JCSCL and its change score were very weakly correlated ($r = 0.16, p < 0.001$), or in other words, city level socio-demographic condition was weakly related to its change during this period.

Table 1 | Correlation coefficients between socio-demographic indicators and time.

| Correlations | Between prefecture average and time | Between city average and time |
|--------------|-----------------------------------|-----------------------------|
| Divorce to marriage ratio | 0.952 | * | 0.961 | ** |
| Percentage elder living alone | 0.999 | *** | 0.999 | *** |
| Percentage nuclear family | −0.868 | † | 0.603 | |
| Percentage living alone | 0.997 | *** | 0.999 | *** |
| Percentage three-generation households | −0.999 | *** | −0.999 | *** |

***p < 0.001, **p < 0.01, *p < 0.05, †p < 0.1.

12The reliability tended to decrease along with time, suggesting more unreliable measurement of JCS in recent years.

13Akita, Toyama, or Tokushima, prefectures that had large population and had neighboring prefectures that were more urban, seem to have changed greatly during the period. In turn, Okinawa and Kagoshima, the most southern prefectures in Japan, showed smallest change. In addition, Tokyo and Kanagawa, two of the most modern prefectures, showed the least change during the period.

14Validity coefficients of the national cultural indicators might greatly attenuate at lower levels such as cities. A similar case was found when national level indicators were collected to examine city level difference in U.K. (0.05 < \alpha < 0.47; Gordon, 2003). However, such low level of reliabilities throughout different years would indicate that either (1) the concept of Individualism-Collectivism (I-C) is not best applicable to describe city differences, or (2) measurement error is unavoidable when locating city level I-C using this specific set of indicators and we need to come up with different indicators at lower levels (i.e., cities).

15For example, the largest change toward individualistic socio-demographic condition took place in Ri-shiri town at the northern island of Hokkaido prefecture, where single living increased from 10% to 40% within two decades. In turn, the largest collectivistic change was observed in A-guni village in Okinawa prefecture, where the percentage of elders living alone decreased from 27.23% to 16.67% within the period.
In Study 1, we measured socio-demographic indicators by prefecture and city, and have observed variation in change. Utilizing this variation as a comparison ground, we conducted Study 2 to examine if Japanese subjective health would differ significantly depending on the degree of change.

**STUDY 2: EXPLAINING ADULT HEALTH FROM CHANGE, PREFECTURE/CITY SOCIO-DEMOGRAPHIC CONDITIONS, AND INDIVIDUAL DIFFERENCES**

We surveyed Japanese adults living in various prefectures in Japan, and tested the hypothesis model ([Figure 2](#)).

**METHOD**

**Participants**

The survey was conducted as a part of a faculty development program to randomly sample graduates of a private university in Hyogo prefecture, one of the prefectures with low score of JCS ([Yamawaki](#), 2012), with students from relatively high familial socio-economic status and high school educational attainment. From the 29,507 graduates who listed their names on the alumni list, 7,628 currently living in Japan were randomly selected as the survey target. They were stratified so as to include every major in the university, and included graduates from every year between 1938 to 2008. Response rate of the survey was 47.2%.

Within the 3,248 collected, we discarded those participants who (1) belonged to a prefecture with less than 10 participants in each prefecture or city, and (2) those who belonged to an outlier city in terms of the distribution of JCSCL, by 2 standard deviations distance from the mean. The selection resulted in 2,343 participants (1617 males, 709 females, 17 unknown, mean age 51.49, $SD = 18.21$) who were living in 53 communities across 17 prefectures.
prefectures (sample distribution is shown in the Supplementary Material, Appendix 1 and 2)\textsuperscript{16}.

\textsuperscript{16}Participants represented the student distribution of the majors in this university (school of theology: 0.3%; literature: 20.0%; sociology: 14.1%; law: 18.4%; economics: 19.6%; commercial science: 19.1%; engineering: 3.1%; policy: 1.8%). The occupation they held the longest during their lifetime varied (labor job: 0.3%; skilled worker: 0.5%; sales person: 16.2%; security staff: 0.3%; office worker: 19.3%; engineer: 3.2%; specialized skill worker: 6.1%; middle management: 21.2%; upper management: 4.9%; self-employed worker: 6.0%; free-lance profession: 2.1%; part-time worker: 2.7%; student: 1.4%; housewife: 3.1%; not employed: 0.2%; other: 2.5%; N/A: 9.7%). Presumably because this was an alumni survey, and the location of the university was in Hyogo, the distribution of the analyzable graduates contained many participants from Osaka ($J$CS $=$ $-5.11$) and Hyogo ($-2.7$) prefectures. Also, the residence of our sample was largely from the individualistic half of the country (Supplementary Material, Appendix 1); therefore, the results should not be understood to cover full range of prefecture variation in Japan. Despite this limitation, a wide age range from working to retired adults who had been working during 1990 to 2010 were obtained, which is valuable for the purpose of this study.

**Measures**

For the explanatory variables, we used the Interdependent Happiness Scale (IHS; Hitokoto and Uchida, 2014), which consists of 9 items asking whether participants experience happiness in their interpersonal relationships. Items included statements such as, “I believe that I and those around me are happy,” or “I feel that I am positively evaluated by others around me.” Participants responded using a 5-point (1: Completely agree—5: Completely disagree), Likert-type scale. The regression weight for this scale score on health measures would respond to our hypothesis 5.

We used the Temperament and Character Inventory (TCI: Cloninger, 1987; Kijima et al., 2000) to measure the personality difference. Because of limited space for questions in the joint survey, we used 2 or 3 representative items that have shown high loadings in the past Japanese study, to tap each of the 4 dimensions of temperament. Examples of TCI items were, “I like to use money rather than save it (novelty seeking),” or “I often get nervous or worried when doing new things I am not familiar with (harm avoidance).” Participants responded using a 4-point (1: Completely agree—4: Completely disagree), Likert-type scale.

We also asked about participants’ annual income, including the annuity, from both the self and the partner as a measure.
of economic condition. Because Japanese adults have large gender differences in working roles, we asked these two separate questions and analyzed by controlling for gender. We asked participants to choose one option from 0: no income, 1:–3.99 million Yen (i.e., approximately less than $40,000), 2: 4–5.99 million Yen ($40,000–60,000), 3: 6–7.99 million Yen ($60,000–80,000), 4: 8–9.99 million Yen ($80,000–100,000), 5: more than 10 million Yen (more than $100,000), separately for themselves and their partner if they were married.

To test our hypothesis 1, we used the General Health Questionnaire (GHQ; Goldberg and Hillier, 1979; Nakagawa and Daibo, 1985). Examples of items from this questionnaire are, “I couldn’t sleep well, because I had worries,” or “I was able to have more fun in my daily life than usual.” Participants rated how they felt/experienced each day using a 4-point (1: I did/didn’t—4: I didn’t/did, depending on the item) Likert-type scale.

Table 2 | Descriptive statistics and reliabilities of the measures/items.

| Measures                      | Min  | Max  | M    | SD   | n of items | α     |
|-------------------------------|------|------|------|------|------------|-------|
| EXPLANATORY VARIABLES        |      |      |      |      |            |       |
| Novelty seeking (TCI<sup>a</sup>) | 3    | 12   | 6.80 | 1.56 | 3          | 0.52  |
| Harm avoidance (TCI)          | 3    | 12   | 8.18 | 1.49 | 3          | 0.58  |
| Reward dependence (TCI)       | 3    | 12   | 8.91 | 1.50 | 3          | 0.64  |
| Persistence (TCI)             | 2    | 8    | 5.57 | 1.15 | 2          | 0.64  |
| Interdependent Happiness (IHS)| 9    | 45   | 33.33| 5.66 | 9          | 0.90  |
| EXPLAINED VARIABLES           |      |      |      |      |            |       |
| General health (GHQ)          | 0    | 12   | 9.56 | 2.84 | 12         | 0.85  |
| Satisfaction with life (SWLS) | 5    | 35   | 22.37| 5.55 | 5          | 0.90  |
| Self-esteem                   | 2    | 10   | 6.96 | 1.37 | 2          | 0.75  |
| Perceived social support in community | 1 | 3    | 0.92 | 0.64 | 1          | –     |
| Perceived social support out of community | 1 | 3    | 1.21 | 0.57 | 1          | –     |

<sup>a</sup> TCI, Temperament and Character Inventory.

To test hypothesis 2, we used the 5-item Satisfaction with Life Scale (SWLS; Diener et al., 1985) as a measure of subjective well-being. The SWLS consisted of items such as, “I am satisfied with my life as a whole,” and participants responded as to whether they would agree to each statement using a 7-point (1: Strongly disagree—7: Strongly disagree) Likert-type scale.

To test hypothesis 3, we used 2 indicator items tapping self-liking (“I like myself”) and self-competence (“I am talented”) aspects of self-esteem (Tafarodi and Swann, 1995; Sakurai, 2000) as a measure of global self-evaluation (Pallant and Lae, 2002; Johnson, 2004). Participants rated how the items applied to themselves using a 5-point (1: Completely agree—4: Completely disagree), Likert-type scale.

To test hypothesis 4, we asked about one’s perceived social support, both in and outside of one’s community, since working and retired adults might have significant others in or outside of the area in which they lived. The question was, “To what extent do you have friends/acquaintances whom you can help or talk to each other?” We asked them to rate separately, “friends/acquaintances in the community (Chi-i-ki in Japanese)” and “friends/acquaintances out of the community,” on a 3-point scale (1: I have many, 2: I have a few, 3: I have none).

Before analysis, scores of GHQ and social support were reversed so that a higher score means better health. For the rest of the scales, higher score indicates better health. Descriptive statistics and reliabilities of the scales are listed in Table 2. Except the four TCI subscales, which showed marginal reliabilities (0.52 < α < 0.64), other measures showed acceptable levels of reliability. Explanatory variables were weakly correlated with each other (Table 3).

Analysis

For each of the dependent measures, we used the same set of explanatory variables (Figure 2). In the model, the first level explanatory variables were, age, gender, four temperaments (i.e., novelty seeking, harm avoidance, reward dependence, and persistence), and interdependent happiness. Second level explanatory variables were the 53 JCSCL scores assigned for...
available communities/cities and the degree of their change from 1990 to 2010. Third level explanatory variables were the 17 JCS scores assigned for available prefectures and the degree of their change from 1990 to 2010. First and second level explanatory variables were group mean centered. The higher level explanatory variables were modeled to explain the intercept of their lower level. Therefore, the model tested would be described as follows.

Level-1 Model:
Health measure (i.e., GHQ) = \( P0 \)  
\[ + P1 \times (Age) + P2 \times (Gender) \]  
\[ + P3 \times (TCI Novelty seeking) \]  
\[ + P4 \times (TCI Harm avoidance) \]  
\[ + P5 \times (TCI Reward dependence) \]  
\[ + P6 \times (TCI Persistence) \]  
\[ + P7 \times (Own income) \]  
\[ + P8 \times (Partner’s income) \]  
\[ + P9 \times (Interdependent happiness) + e \]

Level-2 Model:
\[ P0 = B00 + B01 \times (JCSCL) + B02 \times (JCSCLChange) + r \]

Because each participant was nested within prefectures and cities, we first examined how much variation in our dependent variables would be accounted for by group differences per se using intra-class correlation (ICC). We used the HAD program (Shimizu et al., 2006) to calculate the ICC and conventional F ratio comparing the variance within and between the prefectures for each of the dependent variables, separately for prefecture and city (Supplementary Material, Appendix 3). Given the fact that the typical ICC in applied research is around 0.02-0.22 (Cheung and Au, 2005), there was very little group difference between prefectures or cities. This would indicate that when using self-reported health data among Japanese adults, there are almost negligible differences between prefectures and cities compared to the individual variation within them. With these small effect sizes in mind, we conducted Hierarchical Linear Modeling using HLM7 (Raudenbush and Bryk, 2002).

RESULTS AND DISCUSSION
The results show strong and stable relationships between individual difference variables and health, and small but significant effects of change (Table 4). This indicates that Japanese health is

| Table 4 | Effects of explanatory variables on mental health measures. |
|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|---|
| Explanatory variables | Parameter | General health | Satisfaction with life | Self-esteem | Social support in community | Social support out of community |
| | | B | SE | p | B | SE | p | B | SE | p | B | SE | p |
| **PREFECTURE LEVEL EFFECTS** | | | | | | | | | | | | | | |
| JCS \( a \) | G001 | −0.03 | 0.04 | −0.04 | 0.07 | 0.00 | 0.02 | 0.02 | 0.01 | † | −0.02 | 0.01 | * |
| JCS Change \( b \) | G002 | 0.01 | 0.04 | −0.21 | 0.07 | * | −0.03 | 0.02 | † | 0.00 | 0.01 | −0.02 | 0.01 | ** |
| City level effects | | | | | | | | | | | | | | |
| JCSCL \( c \) | B01 | 0.08 | 0.13 | 0.03 | 0.27 | −0.17 | 0.06 | ** | 0.06 | 0.03 | * | −0.01 | 0.02 | |
| JCSCL Change \( d \) | B02 | −0.05 | 0.08 | −0.02 | 0.16 | −0.04 | 0.03 | −0.02 | 0.02 | 0.00 | 0.02 | | |
| **INDIVIDUAL LEVEL EFFECTS** | | | | | | | | | | | | | | |
| Age | P1 | 0.02 | 0.00 | *** | −0.02 | 0.01 | ** | 0.00 | 0.00 | ** | 0.00 | 0.00 | *** |
| Gender \( e \) | P2 | −0.35 | 0.14 | * | −0.42 | 0.26 | 0.03 | 0.07 | −0.01 | 0.04 | 0.02 | 0.03 | |
| TCI Novelty seeking | P3 | −0.02 | 0.03 | 0.07 | 0.06 | 0.00 | 0.02 | 0.02 | 0.01 | * | 0.02 | 0.01 | * |
| TCI Harm avoidance | P4 | −0.29 | 0.04 | *** | −0.36 | 0.06 | *** | −0.15 | 0.02 | *** | −0.05 | 0.01 | *** | −0.03 | 0.01 | *** |
| TCI Reward dependence | P5 | −0.13 | 0.03 | *** | 0.03 | 0.06 | ** | 0.06 | 0.02 | ** | 0.04 | 0.01 | *** | 0.04 | 0.01 | *** |
| TCI Persistence | P6 | −0.01 | 0.05 | 0.36 | 0.08 | 0.18 | 0.02 | *** | 0.01 | 0.01 | 0.04 | 0.01 | *** |
| Own income | P7 | −0.04 | 0.04 | 0.12 | 0.06 | † | 0.01 | 0.02 | −0.05 | 0.01 | *** | 0.00 | 0.01 | |
| Partner’s income | P8 | −0.04 | 0.04 | 0.16 | 0.07 | * | −0.01 | 0.02 | 0.02 | 0.01 | * | −0.01 | 0.01 | |
| Interdependent Happiness | P9 | 0.23 | 0.01 | *** | 0.59 | 0.02 | *** | 0.10 | 0.00 | *** | 0.02 | 0.00 | *** | 0.01 | 0.00 | *** |

\( a \) Japanese Collectivism Score. Higher the score, the more prefecture level collectivistic socio-demographic condition.

\( b \) Change of Japanese Collectivism Score. Higher the score, the more prefecture change toward individualistic socio-demographic condition.

\( c \) Japanese Collectivism Score at city level. Higher the score, the more city level collectivistic socio-demographic condition.

\( d \) Change of Japanese Collectivism Score at city level. Higher the score, the more city change toward individualistic socio-demographic condition.

\( e \) Gender was coded as 1 = male, 2 = female.

\( TCI \), Temperament and Character Inventory

\(* p < 0.001, ** p < 0.01, * p < 0.05, † p < 0.1 \)
explained largely by individual level factors, and to a lesser extent, by the temporal change of socio-demographic conditions at both prefecture and city levels.

In addition to such a general association, there were differences in the combinations of these societal variables affecting specific health-related variables (Table 4). Contrary to our hypothesis 1, general health was not explained by change, but was explained more by individual age, gender, harm avoidance and reward dependence, and interdependent happiness.

Supporting our hypothesis 2, satisfaction with life was low among those who live in changed prefectures (i.e., negative weight of JCS change on satisfaction with life in Table 4), but city level change did not show the same weight.

Supporting our hypothesis 3 weakly, self-esteem also tended to be low in changed prefectures (i.e., marginally negative weight of JCS change on self-esteem). Negative weight of JCSCL on self-esteem might reflect a relationship between collectivism and low self-esteem (Heine et al., 2001). However, these predicted effects were smaller than the individual differences in harm avoidance, persistence and interdependent happiness, suggesting that subjective well-being and positive self-regard are largely predicted by individual emotional predispositions and interdependent happiness, if within-country difference is concerned.

As for hypothesis 4, social support in one’s community was not explained by change, but tended to be higher among collectivistic socio-demographic prefectures, and high among such communities (i.e., marginally and significantly positive weight of JCS and JCSCL on social support in community). Social support out of one’s community was negatively explained by prefecture change (i.e., negative weight of JCS change on social support out of community), suggesting deterioration of distal support during this period. Collectivistic socio-demographic prefectures, but not cities, showed lower support from outside of one’s community (i.e., negative weight of JCS on social support out of community). Again, individual level factors predicted social support significantly, and harm avoidance and reward dependence, as well as interdependent happiness, were the significant predictors of social support.

Supporting our hypothesis 5, in all health measures examined, interdependent happiness showed positive explanatory weight. This is an indication that interdependent happiness—relational harmony, quiescence, and ordinariness among Japanese adults—is an integral part of their health, independent of change during this period.

**GENERAL DISCUSSION**

In this study, we examined the relationship between socio-demographic change and health, using Japanese adults living in various prefectures and cities. We tested whether the change pertaining to the increase of individualistic socio-demographic condition is related to health during 1990–2010, when the Japanese economy and culture was undergoing massive economic turmoil (Norasakkunkit et al., 2012). As a result, most of our hypotheses were supported and some were not. That is, hypotheses 2 regarding the negative prediction of life satisfaction was supported and hypothesis 3 on self-esteem was weakly supported. Also, hypothesis 5 which stated the positive prediction of interdependent happiness onto health measures was supported. Regarding hypothesis 4 which stated the negative impact of change onto perceived social support, showed mixed results. That is, change at the prefecture level did relate to less perceived social support out of one’s community, but did not relate to that of inside of one’s community. Finally, hypothesis 1 with regard to the negative prediction of change onto general health was not supported. Considering the overall pattern of results, the consequences of economic turmoil might be related to self-evaluative and the social domain of health.

**Pan-cultural implications**

The present findings support the argument that rapid socio-demographic change can be related to psychological difficulties (Norasakkunkit et al., 2012). We may be able to extend such argument by focusing on the socio-demographic change and consider adult health. Also, the result suggests the importance of cultural well-being (Uchida and Ogihara, 2012) as a psychological strength, and basic individual difference in sustaining the health, since the best predictors of health were interdependent happiness and personality. If individual differences are in turn sustained by the larger national culture (Kitayama et al., 2007; Hitokoto and Uchida, 2014), then both personal and social explanations might need to be used together (Iwata, 2006) to explain adult health. Future studies will need to clarify more concrete processes of this contextual aspect of our health, such as how, and what events surrounding the individual would mediate this cross-level impact on health.

**Implications for Japanese culture**

Current results showed that the socio-demographic change which took place during this period was negatively related to the Japanese well-being. Specifically, in more changed prefectures, such as Akita, Tokushima or Toyama prefectures, adults are more unsatisfied with their life, tend to have lower self-esteem, and less social support from out of one’s community compared to adults in other prefectures that underwent smaller amounts of change. Additionally, prefecture and city level collectivistic socio-demographic condition was positively related to social support in one’s community. Our study also resonates with the findings on community strength in Japan (Fukushima et al., 2011), in that both types of research focus on societal level resources above and beyond the individual, that in turn sustain residents’ health. Future studies need to disentangle the correlates between the socio-demographic condition we used and other societal resources in Japan.

**Limitations of the study**

In order to disentangle the relationship between change and health one step further, more effective research design would include collecting cross-sectional data measuring a set of health variables.
and personality variables from the same, as well as age-diverse participants over time. Because our individual data came from a single-shot survey, the results may differ when within-subject data are used to control for participants’ past health, as well as past individual level values or personality. By using a large data set involving cohort design (Ryff, 2008) we would be able to evaluate the effect of change on the deterioration of health in individuals.

In our analytical model we assumed forcefully that individuals had not lived in different prefectures or cities, which is probably not the case. In the joint survey we used in Study 2, we additionally measured individuals’ original prefecture defined as the prefecture participants had lived longest before reaching 15 years old. The results we found for the living prefecture did not emerge when participants are grouped into their original prefecture. Even though Japanese do not relocate as often as Americans on average (Kitayama et al., 2006), residential moves strongly affect how individuals construe the self and positive affective experience (Oishi et al., 2007), regardless of country (Oishi et al., 2009). It is vital for future research to disentangle the factor of residential moves from what we found regarding prefecture and city socioeconomic conditions, and test if our findings are robust after controlling for individual moves. In doing so, it might also be wise to take into account age at the time of the move (Cheung et al., 2010) as well as the specific combinations of the locations of moves.

Further, considering the time span of change in the socioeconomic condition in Japan (Hamamura, 2012), one can also argue that Japanese health problems might have had roots long before the 1990. Locating the temporal cause may be extremely difficult without the guide of any theory, but finding an auto correlation between variables (Varnum and Grossman, 2013), or comparing the same set of variables across different time periods (Oishi et al., 2011) might be an interesting investigation.

It must be noted that our analysis contained prefecture scores that showed negligible ICC. Especially, our city level socio-economic condition showed low reliability, and this may question the reliability of our final outcomes. We note that removing city level factors from the HLM model did not alter our overall findings for other levels. If we were able to find different dimensions that successfully describe cross-regional differences in health, it might override the effect of individual differences we observed in this study. It could also be the case that simply applying national level cultural dimensions to lower level socioeconomic conditions may not be useful to describe health. National level evaluations and their lower counterparts may better be kept separate, so that we can theorize and search for more reliable prefecture, or city, or even lower dimensions of subcultures that provide successful contextual explanations that are open to, therefore more suitable to, describing practical problems and considering remedies.

Finally, we did not model interactions between different levels in our analysis. That is, there still remains a possibility that the explanatory effects of individual level variables on health are systematically different across regional contexts. Such interactions can be hypothesized when we are able to deduce intricate hypotheses regarding correlational differences across regional as well as temporal contexts.

**SUMMARY**

In this study, we started from the research question: what are the psychological consequences of economic stagnation accompanying demographic changes away from the traditional way of life? Based on this, we prepared five hypotheses testing negative relationships between prefecture and city level change toward individualistic socio-demographic condition from 1990 to 2010, with five variables of individual health. The results overall suggest marginal to significant weights of change on life satisfaction, self-esteem, and social support out of someone’s community. We also found positive effects of prefecture or city level JCS (Yamawaki, 2012) on perceived social support in one’s community, as well as large weights of interdependent happiness (Hitokoto and Uchida, 2014) and temperament at the individual level. Future studies should discern further which time period of change or level of socio-demographic condition would truly affect our important life outcomes, using cross-sectional and more representative samples and taking into account the mobile nature of recent Japanese life.

**SUPPLEMENTARY MATERIAL**

The Supplementary Material for this article can be found online at: http://www.frontiersin.org/journal/10.3389/fpsyg.2014.01221/abstract

**REFERENCES**

Allen, M. W., Ng, S. H., Ikeda, K., Jawan, J. A., Sufi, A. H., Wilson, M., et al. (2007). Two decades of change in cultural values and economic development in eight East Asian and Pacific island nations. *J. Cross Cult. Psychol.* 38, 247–269. doi: 10.1177/0022022107300273

Bardi, A., and Goodwin, R. (2011). The dual route to value change: individual processes and cultural moderators. *J. Cross Cult. Psychol.* 42, 271–287. doi: 10.1177/0022022110396916

Baumeister, R. F., and Leary, M. R. (1995). The need to belong: desire for interpersonal attachments as a fundamental human motivation. *Psychol. Bull.* 117, 497–529. doi: 10.1037/0033-2909.117.3.497

Bilsky, W., Janik, M., and Schwartz, S. H. (2011). The structural organization of human values— evidence from three rounds of the European Social Survey (ESS). *J. Cross Cult. Psychol.* 42, 759–776. doi: 10.1177/0022022110362757

Cheung, B. Y., Chudek, M., and Heine, S. J. (2010). Evidence for a sensitive period for acculturation: younger immigrants report acculturating at a faster rate. *Psychol. Sci.* 22, 147–152. doi: 10.1177/0956797610394661

Cheung, M. W. L., and Au, K. (2005). Applications of multilevel structural equation modeling to cross-cultural research. *Struct. Equ. Modeling* 12, 598–619. doi: 10.1207/s15328390sem1204_5

Cloninger, C. R. (1987). A systematic method for clinical description and classification of personality variants: a proposal. *Arch. Gen. Psychiatry* 44, 573–588. doi: 10.1001/archpsyc.1987.01800180093014

Dierer, E. (2000). Subjective well-being: the science of happiness and a proposal for a national index. *Am. Psychol.* 55, 34–43. doi: 10.1037/0003-066X.55.1.34

Diener, E., Emmons, R. A., Larsen, R. J., and Griffin, S. (1985). The satisfaction with life scale. *J. Pers. Assess.* 37, 257–269. doi: 10.1207/s15327752apa4901_13

Draguns, J. G., and Tanaka-Matsumi, J. (2003). Assessment of psychopathology across and within cultures: issues and findings. *Behav. Res. Ther.* 41, 755–776. doi: 10.1016/S0005-7967(02)00190-0

Fukushima, S., Yoshikawa, G., Saizen, I., and Kobayashi, S. (2011). Analysis on the interrelationships between generalized and community trusts: multilevel analysis on all rural areas of three municipalities in northern Kyoto prefecture. *Pop. Env. Inf. Sci.* 25, 137–142. doi: 10.1371/journal.pone.0091975

Goldberg, D. P., and Hillier, V. F. (1979). A scaled version of the general health questionnaire. *R. Soc. Med.* 22, 147–152. doi: 10.1207/s15327752apa4901_13

Gordon, D. (2003). “Area-based deprivation measures: a U.K. perspective,” in *Neighborhoods and Health*, eds I. Kawachi and L. F. Berman (London: Oxford University Press) p. 53.
Hitokoto, H., and Tanaka-Matsumi, J. (2014). Living in the tide of change: explaining Japanese subjective health from the socio-demographic change. Front. Psychol. 5:1221. doi: 10.3389/fpsyg.2014.01221

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