Extremal properties of the variance and the quantum Fisher information

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We show that the variance is its own concave roof. For rank-2 density matrices and operators with zero diagonal elements in the eigenbasis of the density matrix, we prove analytically that the quantum Fisher information is four times the convex roof of the variance. Strong numerical evidence suggests that this statement is true even for operators with nonzero diagonal elements or density matrices with a rank larger than 2. We also find that within the different types of generalized quantum Fisher information considered in [D. Petz, J. Phys. A: Math. Gen. 35, 929 (2002); P. Gibilisco, F. Hiai, and D. Petz, IEEE Trans. Inf. Theory 55, 439 (2009)], after appropriate normalization, the quantum Fisher information is the largest. Hence, we conjecture that the quantum Fisher information is four times the convex roof of the variance even for the general case.

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I. INTRODUCTION

Metrology plays an important role in many areas of physics and engineering. Quantum metrology \cite{1,2} is a novel development, which takes into account the quantum nature of matter when determining the achievable maximal accuracy in parameter estimation. For instance, quantum metrology shows that it is possible to reach a much higher accuracy in typical metrological tasks in a linear interferometer with highly entangled quantum systems than with non-entangled ones, in which no real quantum interaction took place between the particles \cite{3–10}. In practical situations, the influence of noise can reduce considerably the gain in accuracy when applying entangled quantum states for metrology \cite{10}.

Quantum variance and quantum Fisher information are two fundamental notions of quantum metrology. When measuring a quantum mechanical observable $A$, the uncertainty of the measurement outcome is described by the variance defined as

$$\langle (\Delta A)^2 \rangle_\varrho = \langle A^2 \rangle_\varrho - \langle A \rangle_\varrho^2,$$  \hspace{1cm} (1)

where $A$ is a Hermitian operator and $\varrho$ is a density matrix. Connected to this uncertainty, when we would like to estimate a parameter of a quantum evolution based on measuring quantum observables, an uncertainty arises also in this parameter estimation. The most fundamental parameter estimation task is estimating the small parameter $\theta$ of the unitary dynamics $U = \exp(-iA\theta)$ via making measurements on the output state (see Fig. 1).

The accuracy of estimating $\theta$ is bounded from below by the famous Cramér-Rao bound as

$$\Delta \theta \geq \frac{1}{\sqrt{F_{Q}^{BC}[\varrho, A]}},$$ \hspace{1cm} (2)

where the quantum Fisher information is defined as \cite{6–10,11,12,13,14}

$$F_{Q}^{BC}[\varrho, A] = 2 \sum_{i,j} \frac{(\lambda_i - \lambda_j)^2}{\lambda_i + \lambda_j} |A_{ij}|^2.$$ \hspace{1cm} (3)

Here $\lambda_i$ are the eigenvalues of the density matrix and $A_{ij}$ are the matrix elements of the operator $A$ in the eigenbasis of the density matrix.

In Refs. \cite{15–17}, it has been shown that it is possible to define generalized variances and various types of generalized quantum Fisher information. In fact, the two notions are closely connected to each other as for every generalized variance there is a corresponding generalized quantum Fisher information. Thus, the question arises: Are the usual variance, Eq. (1), and the usual quantum Fisher information, Eq. (3), special among the generalized quantities? In this paper we will answer these questions affirmatively. Our findings show that fundamental quantities used in metrology have extremal properties.
In particular, in the first part of the paper we will define generalized variances. The definition presented is broader than that of Refs. [15–17]. Then, we will show that the usual variance is the smallest generalized variance and, connected to this fact, it fulfills the following important relation.

**Theorem 1.** Let us denote the decomposition of a density matrix into the mixture of pure states as

\[ \varrho = \sum_k p_k |\Psi_k\rangle \langle \Psi_k|, \]  

(4)

Then, the variance for a mixed state can be obtained through a concave roof construction as

\[ (\Delta A)^2_\varrho = \sup_{\{p_k, |\Psi_k\rangle\}} \sum_k p_k (\Delta A)^2 \psi_k, \]  

(5)

The proof of Theorem 1 will be given later.

In the second part of the paper we define the generalized quantum Fisher information. The definition presented is also broader than that of Refs. [15–17]. We will show that \( F_{Q}^{BC} \) defined in Eq. (3) is the largest among the types of generalized quantum Fisher information for the low rank case, and connected to this fact, it fulfills the following relation.

**Theorem 2.** For rank-2 density matrices and for an \( A \) with zero diagonal elements in the eigenbasis of the density matrix, the quantum Fisher information can be given as the convex roof of the variance as

\[ F_{Q}^{BC}[\varrho, A] = 4 \inf_{\{p_k, |\Psi_k\rangle\}} \sum_k p_k (\Delta A)^2 \psi_k, \]  

(6)

where \( \{p_k, |\Psi_k\rangle\} \) refers to a decomposition of \( \varrho \) of the type Eq. (3). The proof of Theorem 2 will also be given later.

We will also discuss that numerical calculations suggest that the left-hand side and the right-hand side of Eq. (6) are very close to each other, even when \( A \) has non-zero diagonal elements or the density matrix has a rank larger than 2.

Concerning the quantum Fisher information, we can also prove the following theorem.

**Theorem 3.** \( F_{Q}^{BC}[\varrho, A] \) defined in Eq. (3) is maximal within the various types of quantum Fisher information defined by Refs. [15–17], if they are normalized such that for pure states they equal \( 4(\Delta A)^2 \).

Concerning Theorem 3, it is important to note that all types of quantum Fisher information defined by Refs. [15–17] are convex. Based on Theorem 3 and the strong numerical evidence mentioned above, we formulate the following conjecture.

**Conjecture 1.** The quantum Fisher information \( F_{Q}^{BC}[\varrho, A] \) defined in Eq. (3) is 4 times the convex roof of the variance for density matrices of any rank and any observable \( A \).

The statements of Theorem 1 and Conjecture 1 can be concisely reformulated as follows. For any decomposition \( \{p_k, |\Psi_k\rangle\} \) of the density matrix \( \varrho \) we have

\[ \frac{1}{4} F_{Q}^{BC}[\varrho, A] \leq \sum_k p_k (\Delta A)^2 \psi_k \leq (\Delta A)^2_\varrho, \]  

(7)

where the upper and the lower bounds are both tight.

The paper is organized as follows. In Sec. II, we define generalized variances and determine the minimal generalized variance. InSec. III, we consider the generalized quantum Fisher information for the rank-2 case, and look for the maximal generalized quantum Fisher information. In Sec. IV, we present numerical calculations for density matrices with a higher rank. In Sec. V, we show that \( F_{Q}^{BC} \) is the largest among the types of quantum Fisher information defined by Refs. [15–17]. In Appendix A, we give some details of the semidefinite programs needed for Sec. IV.

II. GENERALIZED VARIANCES AND THE PROOF OF THEOREM 1

In this section, we will define the generalized variance and the minimal generalized variance. We will also show that the minimal generalized variance is the usual variance given in Eq. (1).

**Definition 1.** The generalized variance \( var_\varrho(A) \) is defined by the following two requirements.

1. For pure states, the generalized variance equals the usual variance

\[ var_\varrho(A) = (\Delta A)^2_\varrho. \]  

(8)

2. For mixed states, \( var_\varrho(A) \) is concave in the state.

There are infinitely many generalized variances that fulfill the requirements of Definition 1. Next, we will show that it is possible to identify the smallest one.

**Definition 2.** The minimal generalized variance \( var_\varrho^{\min}(A) \) is defined by the following two requirements.

1. For pure states, it equals the usual variance

\[ var_\varrho^{\min}(A) = (\Delta A)^2_\varrho. \]  

(9)

2. For mixed states, it is defined through a concave roof construction

\[ var_\varrho^{\min}(A) = \sup_{\{p_k, |\Psi_k\rangle\}} \sum_k p_k (\Delta A)^2 \psi_k, \]  

(10)

where \( \{p_k, |\Psi_k\rangle\} \) refers to a decomposition of \( \varrho \) of the type Eq. (3).

Note the counterintuitive fact that while a supremum is used to define Eq. (10), it is the minimal generalized variance and not the maximal one. Let us show that it is indeed the minimal generalized variance. It is clear that for any generalized variance, we have

\[ var_\varrho(A) \geq var_\varrho^{\min}(A). \]  

(11)

If this were not true then it would be possible that for some decomposition of the type Eq. (4) we have...
\( \text{var}_\rho(A) < \sum_k p_k \text{var}_{\Psi_k}(A) \), which would contradict the concavity of the variance. On the other hand, Eq. \([10]\) is concave in the state, since it is defined by a concave roof construction, thus it fulfills Definition 1. Hence, Eq. \([10]\) is the minimal generalized variance.

Next we will prove two lemmas that we need later in the proof of the central theorem of the section.

**Lemma 1.** For rank-2 states, the minimal generalized variance is the usual variance given in Eq. \([1]\). In particular, this statement is true for all qubit states.

**Proof.** For all decompositions of the form Eq. \([1]\), the minimal variance is bounded from above and below as

\[
(\Delta A)^2_{\bar{\rho}} \geq \text{var}^\text{min}_\rho(A) \geq \sum_k p_k (\Delta A)^2_{\Psi_k}.
\]  

(12)

The first inequality in Eq. \([12]\) is due to Eq. \([11]\), the second one is due to the concavity of the variance.

In order to proceed to prove Lemma 1, we need to know an important property of the usual variance \((\Delta A)^2\). It can be written in the following way

\[
(\Delta A)^2_{\bar{\rho}} = \sum_k p_k [(\Delta A)^2_{\Psi_k} + (\langle A \rangle_{\Psi_k} - \langle A \rangle_{\bar{\rho}})^2].
\]

(13)

Here the summation is over the pure subensembles. \((\Delta A)^2_{\Psi_k}\) is the variance in the \(k\)th subensemble, while the \((\langle A \rangle_{\Psi_k} - \langle A \rangle_{\bar{\rho}})^2\) term is the square of the difference between the expectation value for the subensemble and the expectation value for the entire ensemble. One can call the two terms the “quantum” and the “classical” part of the variance, respectively.

Based on Eq. \([13]\) we can state the following. If for a quantum state \(\bar{\rho}\) there is a decomposition \(\{\bar{\rho}_k, \Psi_k\}\) such that the subensemble expectation values equal the expectation value for the entire ensemble (i.e., \(\langle A \rangle_{\Psi_k} = \langle A \rangle_{\bar{\rho}}\) for all \(k\)) then for such a decomposition we have

\[
(\Delta A)^2_{\bar{\rho}} = \sum_k \bar{\rho}_k (\Delta A)^2_{\Psi_k}.
\]

(14)

Based on Eq. \([12]\), we can now state that for \(\bar{\rho}\) the minimum generalized variance of \(A\) and the usual variance are equal

\[
(\Delta A)^2_{\bar{\rho}} = \text{var}^\text{min}_\rho(A).
\]

(15)

We will now show that such a decomposition, in which the subensemble expectation values are all equal, always exists for rank-2 density matrices. Hence, we will show that for all such states the usual quantum variance is the minimal generalized variance.

Let us take an eigendecomposition of the state \(\bar{\rho}\)

\[
\bar{\rho} = p|\Psi_1\rangle\langle \Psi_1| + (1 - p)|\Psi_2\rangle\langle \Psi_2|.
\]

(16)

Thus, in the \(|\Psi_1\rangle\) and \(|\Psi_2\rangle\) basis we can write the density matrix \(\rho\) as

\[
\rho = \begin{bmatrix} p & 0 \\ 0 & 1 - p \end{bmatrix}.
\]

(17)

Let us define now the family of states

\[
|\Psi_\phi\rangle = \sqrt{p}|\Psi_1\rangle + \sqrt{1 - p}|\Psi_2\rangle e^{i\phi}.
\]

(18)

For the states \([18]\), the expectation value of the operator \(A\) can be written as

\[
\langle A \rangle_{\bar{\rho}} = \langle A \rangle_\phi + 2\sqrt{p(1 - p)} \text{Re} \left( \langle \Psi_1 | A | \Psi_2 \rangle e^{i\phi} \right).
\]

(19)

Clearly, there is an angle \(\phi_1\) such that

\[
\text{Re} \left( \langle \Psi_1 | A | \Psi_2 \rangle e^{i\phi_1} \right) = 0.
\]

(20)

For this angle, the expectation value of \(A\) in the state \(|\Psi_{\phi_1}\rangle\) is the same as its expectation value in \(\bar{\rho}\), since the second term in Eq. \([19]\) is zero. In the basis of the states \(|\Psi_1\rangle\) and \(|\Psi_2\rangle\), we can write the projection operators onto \(|\Psi_{\phi_1}\rangle\) as

\[
|\Psi_{\phi_1}\rangle\langle \Psi_{\phi_1}| = \begin{bmatrix} p & \sqrt{p(1 - p)} e^{i\phi_1} \\ \sqrt{p(1 - p)} e^{-i\phi_1} & 1 - p \end{bmatrix}.
\]

(21)

The same is true for \(\phi_1 + \pi\). In the basis of the states \(|\Psi_1\rangle\) and \(|\Psi_2\rangle\), we can write the projection operator onto \(|\Psi_{\phi_1 + \pi}\rangle\) as

\[
|\Psi_{\phi_1 + \pi}\rangle\langle \Psi_{\phi_1 + \pi}| = \begin{bmatrix} p & e^{-i\phi_1} \\ -e^{i\phi_1} & 1 - p \end{bmatrix}.
\]

(22)
Based on Eqs. (21) and (22), the state \( \varrho \) can be decomposed as
\[
\varrho = \frac{1}{2} (|\Psi_{\phi_1}\rangle\langle \Psi_{\phi_1}| + |\Psi_{\phi_1+\pi}\rangle\langle \Psi_{\phi_1+\pi}|),
\]
and the two component states have \( \langle \Psi_{\phi_1}|A|\Psi_{\phi_1}\rangle = \langle A \rangle_\varrho \) and \( \langle \Psi_{\phi_1+\pi}|A|\Psi_{\phi_1+\pi}\rangle = \langle A \rangle_\varrho \), as expected.

In Fig. 2 we show an example to demonstrate how Lemma 1 works for single-qubit states. The three coordinate axes are the expectation values of the three Pauli spin matrices. Points on the surface of the Bloch ball correspond to pure states, points within the ball correspond to rank-2 mixed states. The operator \( A \) has the same expectation value in the states \( \varrho, |\Psi_{\phi_1}\rangle \) and \( |\Psi_{\phi_1+\pi}\rangle \). Note that \( \sigma_z \) has also the same expectation value in these states.

Next we will consider the decomposition of states with a rank larger than 2.

**Lemma 2.** Let us consider an eigendecomposition of a density matrix
\[
\varrho_0 = \sum_{k=1}^{r_0} \lambda_k |\Psi_k\rangle\langle \Psi_k|,
\]
with all \( \lambda_k > 0 \). Let us denote the rank of the density matrix as \( r(\varrho_0) = r_0 \). We consider density matrices for which \( r_0 \geq 3 \). Let us denote by \( A_0 \) the expectation value of an operator \( A \) in the state \( \varrho_0 \)
\[
\text{Tr}(A\varrho_0) = A_0.
\]
We claim that for any \( A \), \( \varrho_0 \) can always be decomposed as
\[
\varrho_0 = p\varrho_- + (1-p)\varrho_+,
\]
such that \( r(\varrho_-) < r_0 \), \( r(\varrho_+) < r_0 \), and
\[
\text{Tr}(A\varrho_+) = \text{Tr}(A\varrho_-) = A_0.
\]

**Proof.** Let us consider the eigenvalues \( \lambda_k \) in the eigendecomposition of \( \varrho \), Eq. (24). First, note that from the conditions \( \text{Tr}(\varrho) = 1 \) and Eq. (25), it follows for the \( \lambda_k \)'s that
\[
\sum_{k=1}^{r_0} \lambda_k = 1,
\]
\[
\sum_{k=1}^{r_0} \lambda_k a_k = A_0,
\]
where we introduced for the subensemble expectation values the notation
\[
a_k = \langle \Psi_k|A|\Psi_k\rangle.
\]
Let us now define a family of states as
\[
\varrho(c, \Delta \vec{\lambda}) = \sum_{k=1}^{r_0} (\lambda_k + c \Delta \lambda_k) |\Psi_k\rangle\langle \Psi_k|,
\]
where \( c \) is a real parameter and \( \Delta \lambda_k \) fulfill
\[
\sum_{k=1}^{r_0} \Delta \lambda_k = 0, \quad \sum_{k=1}^{r_0} a_k \Delta \lambda_k = 0.
\]
Equation (31a) ensures that the state \( \varrho(c, \Delta \vec{\lambda}) \) has a unit trace, while due to Eq. (31b) for all \( c \)
\[
\text{Tr}[A\varrho(c, \Delta \vec{\lambda})] = A_0.
\]
We can rewrite Eqs. (31a) and (31b) as
\[
\begin{bmatrix}
a_1 & a_2 & \ldots & a_{r_0-1} & a_{r_0} \\
1 & 1 & \ldots & 1 & 1
\end{bmatrix} \Delta \vec{\lambda} = \begin{bmatrix} 0 \\
0 \end{bmatrix}. \tag{33}
\]
For \( r_0 \geq 3 \), there is always a \( \Delta \vec{\lambda} \neq 0 \) fulfilling Eq. (33). It is clear that, if \( \varrho(c, \Delta \vec{\lambda}) \) has non-negative eigenvalues, then \( \varrho(c, \Delta \vec{\lambda}) \) is a physical state. In the following, for compactness, we will omit \( \Delta \vec{\lambda} \) from the argument of \( \varrho \).

Let us now examine the properties of the \( \varrho(c) \) family. It is clear that
\[
\varrho(0) = \varrho_0.
\]
Moreover, it is also clear that around \( c = 0 \) there is some range of values for \( c \), such that \( \varrho(c) \) is physical. So first, let us start to increase \( c \) gradually from zero until one of the eigenvalues becomes zero. Let us denote this value by \( c_+ \). Thus, \( \varrho(c) \) is physical for
\[
0 \leq c \leq c_+.
\]
We have \( r(\varrho(c_+)) < r_0 \). Moreover, \( \varrho(c) \) is non-physical for \( c > c_+ \). A similar thing happens, if we start from \( c = 0 \) by decreasing \( c \) gradually until one of the eigenvalues becomes zero. Let us denote this value by \( c_- \). Thus, \( \varrho(c) \) is physical for
\[
c_- \leq c \leq 0.
\]
Again, we have \( r(\varrho(c_-)) < r_0 \). Moreover, \( \varrho(c) \) is non-physical for \( c < c_- \). In summary, for any choice of \( \varrho \) and \( \Delta \lambda_k \) fulfilling Eqs. (31b) and (31a), there is a family of states \( \varrho(c) \) defined in Eq. (30) such that \( \varrho(c) \) is physical for
\[
c_- \leq c \leq c_+,
\]
and \( r(\varrho(c_+)) < r_0 \), and \( r(\varrho(c_-)) < r_0 \). One can explicitly reconstruct these bounds as
\[
c_+ = \min_k \frac{\lambda_k}{\Theta(-\Delta \lambda_k)} \tag{38}
\]
and
\[
c_- = -\min_k \frac{\lambda_k}{\Theta(\Delta \lambda_k)}, \tag{39}
\]
FIG. 3: (Color online) The rank-3 mixed state $\rho_0$ is decomposed as Eq. (26) into the mixture of two rank-2 states, $\rho_-$ and $\rho_+$. The coordinate axes are the eigenvalues of the density matrix, while we assume that all density matrices have the same eigenstates as $\rho$. Points corresponding to states $\rho'$ for which $\text{Tr}(\rho' A) = \text{Tr}(\rho_0 A)$ are on the dashed line. For $\rho_+$ we have $\lambda_2 = 0$, while for $\rho_-$ we have $\lambda_1 = 0$.

where

$$\Theta(x) = \begin{cases} x & \text{if } x \geq 0, \\ 0 & \text{if } x < 0. \end{cases}$$ (40)

Let us now construct the decomposition Eq. (26). Direct calculations can show that the following decomposition is of the desired form

$$\rho_- = \rho(c_-),$$ (41)

and

$$\rho_+ = \rho(c_+),$$ (42)

while the mixing probability is

$$p = \frac{c_+}{c_+ - c_-}. \quad (43)$$

In Fig. 3 we presented an example showing how Lemma 2 works for a rank-3 state. A point $(\lambda_1, \lambda_2, \lambda_3)$ in the coordinate system corresponds to $\rho = \sum_{k=1}^3 \lambda_k |\Psi_k\rangle \langle \Psi_k|$. All points have to fulfill Eq. (28) and $\lambda_k \geq 0$ for $k = 1, 2, 3$. Thus, all points corresponding to physical states are in the triangle. Interior points of the triangle correspond to rank-3 states, while points on the boundary correspond to rank-2 and rank-1 states.

After proving Lemma 1 and Lemma 2, we will prove the main theorem of the section.

Proof of Theorem 1. Theorem 1 is equivalent to the statement that for systems of any dimension, the minimal generalized variance is the usual variance

$$\text{var}_A^{\text{min}}(A) = (\Delta A)_g^2. \quad (44)$$

For rank-2 states, this is true because of Lemma 1. According to Lemma 2, any state with a rank larger than 2 can be decomposed into the mixture of lower rank states that have the same expectation value for $A$ as the original state has. The lower rank states can then be decomposed into the mixture of states with an even lower rank, until we reach rank-2 states. Thus, any state $\rho$ can be decomposed into the mixture of states with an even lower rank, until we reach rank-2 states. Thus, any state $\rho$ can be decomposed into the mixture of the form Eq. (4) such that $\langle \Psi_k| A |\Psi_k\rangle = \text{Tr}(A \rho)$. Hence, similarly ro the proof of Lemma 1, the statement of Theorem 1 follows [20]. □

We will present two examples in order to clarify certain issues concerning generalized variances.

Example 1. Let us consider the state

$$\rho = \frac{1}{2} \left( |+1\rangle \langle +1| + |-1\rangle \langle -1| \right), \quad (45)$$

where $| \pm 1 \rangle$ denote the eigenstates of the Pauli spin matrix $\sigma_z$. For this state, we have $(\Delta \sigma_z)^2 = +1$. Let us look at the decomposition

1. $p_1 = \frac{1}{2}, |\Psi_1\rangle = \frac{1}{\sqrt{2}}(| +1 \rangle + | -1 \rangle), \quad (46)$
2. $p_2 = \frac{1}{2}, |\Psi_2\rangle = \frac{1}{\sqrt{2}}(| +1 \rangle - | -1 \rangle).$

For this decomposition $(\Delta \sigma_z)^2 p_k = +1$ and $\langle \Psi_k| \sigma_z |\Psi_k\rangle = \text{Tr}(\sigma_z \rho) = 0$. Thus, we presented a decomposition such that the expectation value of $A$ is the same for both subensembles $|\Psi_k\rangle$.

Example 2. Let us consider the generalized variance

$$\text{var}^{\text{quadratic}}_A(\rho) = \text{Tr}(A^2 \rho) - \text{Tr}(A \rho)^2 + 1 - \text{Tr}(\rho^2). \quad (47)$$

For pure states, it equals the usual variance. For mixed states, it is larger than that and it is concave in the state. It is in fact a sum of the usual variance and the linear entropy, which is zero for pure states and is concave in the state.

III. GENERALIZED QUANTUM FISHER INFORMATION AND THE PROOF OF THEOREM 2

In this section, we will define the generalized quantum Fisher information and the maximal generalized quantum Fisher information. We will show that for rank-2 density matrices, the maximal quantum Fisher information is the usual quantum Fisher information.

Definition 3. The generalized quantum Fisher information $F_Q[\rho, A]$ is defined by the following two requirements.

1. For pure states, we have

$$F_Q[|\Psi\rangle \langle \Psi|, A] = 4(\Delta A_\rho)^2.$$

The factor 4 appears to keep the consistency with the existing literature [6].
2. For mixed states, $F_Q[\varrho, A]$ is convex in the state.

The usual quantum Fisher information, Eq. (3), is one of the types of generalized quantum Fisher information. Next, we will present another well-known generalized quantum Fisher information.

**Example 3.** Let us consider the Wigner-Yanase skew information defined as [21]

$$I[\varrho, A] = \text{Tr}(A^2 \varrho) - \text{Tr}(A \varrho A \varrho) .$$

(48)

For pure states it equals $(\Delta A)^2$ and it is convex in the state. Thus $4I[\varrho, A]$ fulfills Definition 3.

There are infinitely many types of generalized quantum Fisher information that fulfill the requirements of Definition 3. Next, we will show that it is possible to identify the largest one.

**Definition 4.** The maximal generalized quantum Fisher information, $F_Q^{\text{max}}[\varrho, A]$, is defined by the following two requirements.

1. For pure states, it equals four times the usual variance

$$F_Q^{\text{max}}[|\Psi\rangle\langle\Psi|, A] = 4(\Delta A)^2 .$$

(49)

2. For mixed states, it is defined through a **convex roof** construction [18]

$$F_Q^{\text{max}}[\varrho, A] = 4 \inf_{\{p_k, |\psi_k\rangle\}} \sum_k p_k (\Delta A)^2 |\psi_k\rangle .$$

(50)

It is clear that for any generalized quantum Fisher information, we have

$$F_Q[\varrho, A] \leq F_Q^{\text{max}}[\varrho, A] .$$

(51)

This can be proven similarly as it has been shown that the minimal generalized variance is smaller than or equal to all other generalized variances.

Next we will present the main theorem of this section. **Proof of Theorem 2.** Let us consider a density matrix $\varrho$ of rank 2 given as

$$\varrho = q|\phi_1\rangle\langle\phi_1| + (1 - q)|\phi_2\rangle\langle\phi_2| ,$$

(52)

and a self-adjoint observable $A$. We will show that for any $\varrho$ and $A$ exists a decomposition into a mixture of projectors of the form Eq. (4) such that

$$F_Q^{\text{BC}}[\varrho, A] = 4 \sum_k p_k (\Delta A)^2 |\psi_k\rangle .$$

(53)

Knowing that

$$\sum_k p_k 4(\Delta A)^2 |\psi_k\rangle \geq F_Q^{\text{max}}[\varrho, A] \geq F_Q^{\text{BC}}[\varrho, A] ,$$

(54)

if a decomposition of the type Eq. (53) exists then

$$F_Q^{\text{BC}}[\varrho, A] = F_Q^{\text{max}}[\varrho, A] .$$

(55)

Note that the first inequality in Eq. (53) is due to the convexity of the quantum Fisher information, the second one is due to Definition 4.

We consider the case where the diagonal elements of $A$ are zero, hence $A$ can explicitly be obtained as

$$A = |A_{12}| \left( |\phi_1\rangle\langle\phi_2| e^{i\alpha} + |\phi_2\rangle\langle\phi_1| e^{-i\alpha} \right)$$

(56)

and $|A_{12}|$ and $\alpha$ are real. In fact, the the phase $\alpha$ is irrelevant. Thus, in our computation only $|A_{12}|$ is essential, so we can reduce the problem to the two-dimensional space generated by $|\phi_1\rangle$ and $|\phi_2\rangle$, and instead of $A$, it is sufficient to take into account $|A_{12}|$. For the usual quantum Fisher information, Eq. (3), we have

$$F_Q^{\text{BC}}[\varrho, A] = 4(2q - 1)^2 |A_{12}|^2 .$$

(57)

For the right-hand side of Eq. (53) we have

$$4 \sum_k p_k (\Delta A)^2 |\psi_k\rangle = 4 \left[ \text{Tr}(\varrho A^2) - \sum_k p_k |\langle \Psi_k | A | \Psi_k \rangle|^2 \right] .$$

(58)

Based on Eqs. (57) and (58), Eq. (53) can be rewritten as

$$4(2q - 1)^2 |A_{12}|^2 = 4 \left[ \text{Tr}(\varrho A^2) - \sum_k p_k |\langle \Psi_k | A | \Psi_k \rangle|^2 \right] .$$

(59)

This is the equality to be solved. Using that

$$A^2 = |A_{12}|^2 \left( |\phi_1\rangle\langle\phi_1| + |\phi_2\rangle\langle\phi_2| \right) ,$$

(60)

we obtain $\text{Tr}(\varrho A^2) = |A_{12}|^2$. Substituting this into Eq. (59), we get

$$4q(1 - q)|A_{12}|^2 = \sum_k p_k |\langle \Psi_k | A | \Psi_k \rangle|^2 .$$

(61)

Finally, we can present a desired decomposition of the form Eq. (4) for $\varrho$ fulfilling Eq. (61). Its is given by two subensembles as

$$p_1 = \frac{1}{2}, \quad |\Psi_1\rangle = \sqrt{q}|\phi_1\rangle + \sqrt{1 - q}|\phi_2\rangle ,$$

(62)

and

$$p_2 = \frac{1}{2}, \quad |\Psi_2\rangle = \sqrt{q}|\phi_1\rangle - \sqrt{1 - q}|\phi_2\rangle .$$

(63)

**IV. NUMERICAL EVIDENCE FOR THE VALIDITY OF THEOREM 2 FOR MATRICES WITH A RANK LARGER THAN 2**

After presenting a proof for Eq. (61) for the rank-2 case, in this section we examine numerical evidence for larger systems. Thus, our approach is similar to the one followed in Ref. [19], where the concurrence formula has been derived for rank-2 mixed states and for mixed states
with a higher rank it has been supported by numerical
evidence.

We show a method to compute the infimum over con-
 vex decompositions in Eq. (50) by mapping this problem
to an optimization over symmetric separable states.

While the optimization for symmetric separable states
can be computed only for \( d = 2 \), we get a lower bound on
the infimum by optimizing for symmetric quantum states
with a positive partial transpose (PPT), since states with a
positive partial transpose are a superset of separable
states. We will also consider a smaller superset, the
a positive partial transpose are a superset of separable
states \([22]\). Note that the optimization for density matrices and
states \([23]\). We will also consider a smaller superset, the
set of quantum states with a PPT symmetric extension
\([24]\). Note that the optimization for density matrices and
operators of dimension \( d \) is equivalent to optimization over density matrices of rank-\( d \) allowing for density matrices and operators of any size.

The optimization problem for \( F_Q^{\text{max}}[\varrho, A] \) defined in
Eq. (50) can be rewritten as follows.

\[
F_Q^{\text{max}}[\varrho, A] = 4 \left( \langle A^2 \rangle_\varrho - \sup_{\{p_k, |\Psi_k\rangle\}} \sum_k p_k \langle A \rangle_{\Psi_k}^2 \right),
\]

where \( \{p_k, |\Psi_k\rangle\} \) refers to a decomposition of \( \varrho \) of
the type Eq. (4).

Next, we will rewrite the term quadratic in expectation
values as a term linear in expectation values of an
operator acting on a bipartite system as

\[
F_Q^{\text{max}}[\varrho, A] = 4 \left( \langle A^2 \rangle_\varrho - \sup_{\{p_k, |\Psi_k\rangle\}} \sum_k p_k \langle A \otimes A \rangle_{\Psi_k \otimes \Psi_k} \right).
\]

Further transformations lead to a form in which instead
of a sum of expectation values we have an expectation
value of a mixture as

\[
F_Q^{\text{max}}[\varrho, A] = 4 \left( \langle A^2 \rangle_\varrho - \sup_{\{p_k, |\Psi_k\rangle\}} \langle A \otimes A \rangle_{\sum_k p_k |\Psi_k\rangle \otimes |\Psi_k\rangle} \right).
\]

Hence, we arrive at a form based on an optimization over
symmetric separable states

\[
F_Q^{\text{max}}[\varrho, A] = 4 \left( \langle A^2 \rangle_\varrho - \sup_{\varrho_s \in S_s, \text{Tr}_1(\varrho_s)=\varrho} \langle A \otimes A \rangle_{\varrho_s} \right)
= 2 \inf_{\varrho_s \in S_s, \text{Tr}_1(\varrho_s)=\varrho} \langle (A \otimes 1 - 1 \otimes A)^2 \rangle_{\varrho_s},
\]

where \( S_s \) is the set of symmetric separable states. States
in \( S_s \) are mixtures of symmetric product states \([22, 25]\)

\[
\sum_k p_k |\Psi_k\rangle \otimes |\Psi_k\rangle^2.
\]

It is not possible to compute the right-hand side of
Eq. (67) in general. However, it is possible to obtain a
series of lower bounds on it based on the theory of PPT
symmetric extensions \([24]\). Let us introduce the notation for
such bounds

\[
\mathcal{B}_{\text{SE}_n}(A, \varrho) := 2 \inf_{\varrho_s \in \mathcal{S}_{\text{SE}_n}} \langle (A \otimes 1 - 1 \otimes A)^2 \rangle_{\varrho_s}, \quad \text{Tr}_1(\varrho_s) = \varrho
\]

(69)

where \( \mathcal{S}_{\text{SE}_n} \) are the symmetric bipartite states with a
PPT symmetric extension of size \( n \). We note that \( \mathcal{S}_{\text{SE}_2} \)
is identical to the set of symmetric PPT states

\[
\mathcal{S}_{\text{SE}_2} \equiv \mathcal{S}_{\text{PPT}}.
\]

(70)

hence we will use the two notations interchangeably in
the rest of the paper. With the definition Eq. (69), for
\( n > 3 \) we can write

\[
F_Q^{\text{max}}[\varrho, A] \geq \ldots \geq \mathcal{B}_{\text{SE}_{(n+1)}}(A, \varrho) \geq \mathcal{B}_{\text{SE}_n}(A, \varrho) \geq \ldots \geq \mathcal{B}_{\text{PPT}}(A, \varrho).
\]

(71)

All these bounds can be obtained using semidefinite pro-
gramming \([24]\).

The non-increasing series of bounds in Eq. (71) exists
since, in general, states having a PPT symmetric exten-
sion for \( n \) qudits are a subset of states having a PPT
symmetric extension for \( m < n \) qudits. In entanglement
theory PPT states are typically separable and entangled
states with a PPT property are very special \([23]\). Thus,
the bound \( \mathcal{B}_{\text{PPT}} \) based on an optimization over PPT
states is expected to be very close to the value obtained
from optimization over separable states for most \( A \)'s and
\( \varrho \)'s. The bound \( \mathcal{B}_{\text{SE}_3} \) computed based on an opti-
mization over \( \mathcal{S}_{\text{SE}_3} \) is expected to be even closer to the one
obtained from an optimization over separable states.

We carried out numerical tests using random matrices.
We considered 10000 random trials for dimen-
sions \( d = 2, 3, 4 \) for computing the bound \( \mathcal{B}_{\text{PPT}} \) defined
in Eq. (69). The matrix \( A \) was generated as follows.
First a matrix \( M \) was obtained such that all the real
and imaginary parts of its elements were independent random
numbers with a normal distribution centered around 0 and
with a variance \( \sigma = 1 \). Then, the Hermitian matrix
\( A \) was obtained from \( M \) as \( \frac{1}{2}(M + M^\dagger) \). The density
matrix \( \varrho \) was generated based on the method described in
Ref. \([26]\). For each randomly chosen \( A \) and \( \varrho \) we det-
\emph{ined the relative difference between \( \mathcal{B}_{\text{PPT}}(A, \varrho) \) given
in Eq. (69) and \( F_Q^{\text{BC}}[\varrho, A] \). We used the program pack-
ages Sedumi and YALMIP for semidefinite program-
ming, while QUBIT4MATLAB V4.0 was used for calcu-
lations connected to quantum physics \([27, 29]\). MATLAB
uses double precision arithmetic.

To test our method, we considered first the \( d = 2 \) case
for \( A \) matrices with zero diagonal elements. According to
Theorem 2 and knowing that the set of PPT states and
the set of separable states are the same for two qubits,
TABLE I: (Top line) Statistics of the relative difference between the quantum Fisher information and $B_{\text{SPPT}}(\rho, A)$ for 10000 random tests. For each test, a random $A$ with zero diagonal elements and a random $\rho$ is generated. The largest relative difference, the average relative difference and the standard deviation of the relative difference (i.e., the square root of the variance) is shown for two qudits for $d = 2$. SPPT indicates that the optimization is taken over PPT symmetric states. (Second, third and fourth lines) The same for random $A$ with nonzero diagonal elements. The same for $d = 3, 4$ for the relative difference between the quantum Fisher information and $B_{\text{SE}_3}(\rho, A)$. $\text{SE}_3$ indicates that the optimization is taken over symmetric states with a PPT symmetric extension for $N = 3$ qudits.

| Task               | Largest Rel. Diff. | Average Rel. Diff. | Standard Dev. of Rel. Diff. |
|--------------------|--------------------|--------------------|-----------------------------|
| 2SPPT, zeros in the diagonal | $1.4793 \times 10^{-6}$ | $3.6194 \times 10^{-10}$ | $2.2321 \times 10^{-8}$ |
| 2SPPT              | $9.2505 \times 10^{-6}$ | $1.3313 \times 10^{-9}$ | $9.8533 \times 10^{-9}$ |
| 3SPPT              | $2.3552 \times 10^{-8}$ | $8.8486 \times 10^{-10}$ | $9.9692 \times 10^{-10}$ |
| 4SPPT              | $9.0345 \times 10^{-9}$ | $1.3448 \times 10^{-9}$ | $9.8216 \times 10^{-10}$ |
| 3SE$_3$            | $7.7749 \times 10^{-8}$ | $4.2694 \times 10^{-10}$ | $3.5161 \times 10^{-9}$ |
| 4SE$_3$            | $9.5506 \times 10^{-9}$ | $1.0429 \times 10^{-9}$ | $1.0404 \times 10^{-9}$ |

The variance according to Refs. [15, 16] is equal to the bound $B_{\text{SPPT}}(\rho, A)$. The results are in the first row of Table I. We found that the largest relative difference is below $2 \times 10^{-6}$, while the average relative difference and the standard deviation of the relative difference are much smaller. The nonzero value must be the result of the finite precision of the numerical calculations.

After testing the method for a case that have been proven analytically in Theorem 2, we also used it to explore the cases that have not been proven yet. We carried out calculations for $A$ matrices with nonzero diagonal elements for $d = 2, 3, 4$. The results of the numerical tests can be seen in the second, third and fourth rows of Table I. We also carried out similar numerical tests for dimensions $d = 3, 4$ for computing the bound $B_{\text{SE}_3}$ based on an optimization for symmetric states with an $N = 3$ PPT symmetric extension given in Eq. (69) and compare it to the quantum Fisher information. The results of the numerical tests can be seen in the bottom part of Table I.

We can see that for all these trials, the relative difference between the two bounds and the value of $F^B_{\text{Q}}(\rho, A)$ is smaller than $10^{-5}$, while the average relative difference is smaller than $10^{-8}$, and the standard deviation of the relative difference is smaller than $10^{-7}$. Thus, the numerical tests suggest that it should be examined carefully through analytical calculations whether

$$F^\text{max}_{\text{Q}}[\rho, A] = F^B_{\text{Q}}[\rho, A]$$

is true in general. Based on this strong evidence, we conjecture that this is the case. The semidefinite programs necessary to calculate $B_{\text{SPPT}}(\rho, A)$ and $B_{\text{SE}_3}(\rho, A)$ are outlined in Appendix A.

V. MAXIMAL QUANTUM FISHER INFORMATION WITHIN THE SET DEFINED BY REF. [15, 16]

In this section, first we will review how Refs. [15, 16] define the variances and the types of generalized quantum Fisher information. We will denote them by $\text{vâr}$ and $\hat{f}$, in order to distinguish them from the generalized variances and various types of Fisher information defined in this paper. Then, we will show that these definitions, apart from a constant factor, fulfill Definition 1 and Definition 3. This is a surprise since these quantities in Refs. [15, 16] were defined based on ideas very different from the ones presented in this paper. We also show that the set of variances and the types of quantum Fisher information satisfying Definition 1 and Definition 3 are a broader set than the corresponding quantities presented in Refs. [15, 16]. Finally, we will show that the usual quantum Fisher information given in Eq. (3), after appropriate normalization, is the largest even within the various types of quantum Fisher information considered in Refs. [15, 16].

A. Summary of the basic results of Refs. [15, 16]

The basic idea of Refs. [15, 16] is that for each standard matrix monotone function $f : \mathbb{R}^+ \to \mathbb{R}^+$, a generalized variance and a corresponding quantum Fisher information are defined. The notation standard means that $f(1) = 1$ and $f(t) = tf(t^{-1})$. First, let us consider the generalized variances.

Definition 5. The variance according to Refs. [15, 16] is defined as

$$\text{vâr}_f(A) = \langle A, J^f(A) \rangle - (\text{Tr} \rho A)^2,$$  \hspace{1cm} (73)

where $A$ is Hermitian, the scalar product is defined as $\langle A, B \rangle = \text{Tr}(AB)$, and

$$J^f(A) = f(L_{\rho} \mathbb{R}_e^{-1}) \mathbb{R}_e,$$  \hspace{1cm} (74)

where

$$L_{\rho}(A) = \rho A, \quad \mathbb{R}_e(A) = A \rho.$$
Computing $\mathcal{J}_\varrho(A)$ can be simplified knowing that for $f(x) = \sum_k c_k x^k$ we have [15, 16]

$$
\mathcal{J}_\varrho(f)(A) = \sum_k c_k \varrho^k A f(1-k)
$$

Thus, Eq. (75) can be used to calculate $\mathcal{J}_\varrho(A)$ for any polynomial of $x$.

Moreover, it is also useful to define the mean based on $f$ as

$$
m_f(a, b) = af\left(\frac{b}{a}\right)
$$

and use it instead of $f$. The $f(1)$ = 1 normalization condition corresponds to the condition $m_f(a, a) = a$ for the means. The $f(t) = tf(t^{-1})$ requirement corresponds to $m_f(a, b) = m_f(b, a)$. A list of quantum Fisher information generated by various well-known means $m_f(a, b)$ can be found in Refs. [15–17].

After we discussed the generalized variances, we will turn our attention to the quantum Fisher information.

**Definition 6.** The quantum Fisher information according to Refs. [15, 16] is defined as

$$
\hat{F}_Q^f(\varrho; A) = \text{Tr}(A \mathcal{J}_\varrho^{-1}(\varrho) A)
$$

(77)

Again, the most important types of generalized quantum Fisher information correspond to famous means appearing often in mathematics.

For the arithmetic mean $m_f(a, b) = \frac{a+b}{2}$, the quantum Fisher information is defined as

$$
\hat{F}_Q^{BC}(\varrho; A) = \int_0^\infty \text{Tr}[\exp(-t\varrho/2)A \exp(-t\varrho/2)A]\text{d}t.
$$

(78)

Other means generate other types of quantum Fisher information. Details can be found in Refs. [15, 17].

**B. The usual quantum Fisher information is the maximal one within the set of quantum Fisher information defined in Refs. [15, 16]**

After reviewing the results of Refs. [15, 16], we will connect the results of the present paper to that of Refs. [15, 16. In particular, we will show that the usual quantum Fisher information, Eq. (3), is the largest within the set of quantum Fisher informations defined by Refs. [15, 16], if they are normalized such that for pure states they equal four times the usual variance given in Eq. (7).

In order to proceed, we will now prove the following lemma concerning the behavior of the family of variances defined in Refs. [15, 16] for pure states.

**Lemma 3** For pure states we have

$$
\text{var}_\varrho^f(A) = 2m_f(1, 0) \times (\Delta A)^2
$$

(79)

Thus, for pure states the variance of Refs. [15, 16] equals the usual variance times a constant.

Proof. It can be shown that Eq. (73) can be rewritten as [16]

$$
\text{var}_\varrho^f(A) = \sum_{ij} m_f(\lambda_i, \lambda_j) |A_{ij}|^2 - |\sum_{ij} \lambda_i A_{ii}|^2
$$

(80)

where $\lambda_j$ are the eigenvalues of $\rho$, and $A$ is given in the basis of the eigenvectors of the density matrix. Here we took advantage of the fact that $m_f(a, a) = 1$. One can see that for pure states ($\lambda_1 = 1, \lambda_k = 0$ for $k \geq 2$) we have

$$
\text{var}_\varrho^f(A) = 2 \sum_{j>1} m_f(1, 0) |A_{1j}|^2.
$$

(81)

Hence, Eq. (79) follows. Note that there is a factor of 2 in Eq. (81) since $m_f(\lambda_i, \lambda_j)$ in Eq. (80) is nonzero for the pure state considered for $i = j$, $j > 1$ and for $j = 1, i > 1$.

Next, we will turn our attention to the quantum Fisher information. We define the following notation

$$
\hat{F}_Q^f(\varrho; A) = \hat{F}_Q^f(\varrho; i[\varrho; A])
$$

(82)

With this definition, for the arithmetic mean $m_f(a, b) = \frac{a+b}{2}$, $\hat{F}_Q^f(\varrho; A)$ equals the usual quantum Fisher information given in Eq. (3) denoted in the physics literature as $F_Q(\varrho; A)$ [6].

It is instructive to analyze the differences between the quantum information $\hat{F}_Q^f(\varrho; A)$ appearing in Refs. [15, 16] and, in general, in the mathematics literature and $F_Q(\varrho; A)$, which appears in the physics literature. The difference between the two notions is that $1/\hat{F}_Q^f(\varrho; A)$ bounds the squared uncertainty when the small parameter $t$ of

$$
\varrho_{\text{output}}(t) = \varrho + At
$$

(83)

is estimated [15]. Here $A$ is a matrix, and hence in this case the output density matrix is a linear function of $t$. In contrast, $1/F_Q(\varrho; A)$ bounds the squared uncertainty when estimating $\theta$ in

$$
\varrho_{\text{output}}(t) = \exp(-iA\theta)\varrho\exp(+iA\theta),
$$

(84)

as mentioned in the introduction. The definition of $F_Q(\varrho; A)$ is not clear for pure states, as for such states $\varrho + At$ is a physical for $t > 0$ or $t < 0$, while Eq. (84) describes a physical dynamics that leads to a valid density matrix for all $t$.

Next, we will prove a connection between the generalized quantum Fisher information and the usual variance for pure states.

**Lemma 4.** For pure states, the generalized quantum Fisher information $\hat{F}_Q^f(\varrho; A)$ is proportional to the usual variance given in Eq. (7)

$$
\hat{F}_Q^f(\varrho; A) = \frac{2(\Delta A)^2}{m_f(1, 0)}.
$$

(85)
fact that $\var\rho$ based on a single-variable function $f$, cannot be written in the form Eq. (73) remembering that $\var$ is a generalized variance according to Definition 1, but equivalent? We find that this is not the case. Eq. (46) this point the question arises: Are the the two definitions 

For $m_f(a, b) = \frac{a+b}{2}$, we obtain the usual quantum Fisher information given in Eq. (3). The lemma can be proven substituting $\lambda_1 = 1$, $\lambda_k = 0$ for $k \geq 2$ into Eq. (57). □

Next, we will present the central theorems of this section.

**Theorem 4.** After appropriate normalization, the generalized variances defined in Refs. [15, 16] fulfill Definition 1. However, the opposite is not true. Not all generalized variances that fulfill Definition 1 belong to the generalized variances defined in Refs. [15, 16].

Proof. It has been proven that the generalized variances of Refs. [15, 16] are concave [30]. Based on this fact and Lemma 3, we see that

$$\var_f(A) = \frac{\var_f(A)}{2m_f(1, 0)}$$

is a generalized variance in the sense of Definition 1. At this point the question arises: Are the the two definitions equivalent? We find that this is not the case. Eq. (46) is a generalized variance according to Definition 1, but cannot be written in the form Eq. (73) remembering that $\var_f(A)$ is defined in Eq. (46). This is essentially due to the fact that $\var^2$ appears in the definition Eq. (46). On the other hand, Refs. [15, 16] defines a family of variances based on a single-variable function $f$, while Definition 1 does not provide an explicit formula for obtaining such quantities.

**Theorem 5.** After appropriate normalization, the types of quantum Fisher information defined in Refs. [15, 16] fulfill Definition 3.

Proof. It has been proven that the types of generalized quantum Fisher information of Refs. [15, 16] are convex [30]. Based on Lemma 4, we can also see that

$$F_Q^f[\var, A] = 2m_f(1, 0)\hat{F}_Q^f[A, A]$$

is a generalized quantum Fisher information in the sense of Definition 3.

Next, we discuss, which quantum Fisher information is the largest from the family considered in Refs. [15, 16]. It has been proven that $\hat{F}_Q^f[A, A]$ is the smallest for $f(x) = \frac{1+x}{2}$, that is for the usual quantum Fisher information [15, 16]. Moreover, it has also been found that the quantum Fisher information for $f(x) = \frac{2x}{1+x}$, i.e., for the harmonic mean, is the smallest [15, 16]. This can be seen as follows. In the expression for the quantum Fisher information, Eq. (86), appears the $m_f(a, b)$ mean. The arithmetic mean $m_f(a, b) = \frac{a+b}{2}$ is the largest among the means, while the harmonic mean is the smallest.

With a similar argument, we can prove that $F^f_Q[\var, A]$ defined in Eq. (87) is the largest for $f(x) = \frac{1+x}{2}$, i.e., for the arithmetic mean. It can also be proven that it is the smallest for the harmonic mean.

According to Theorem 4, the variance fitting Definition 1 is of the form Eq. (88). Moreover, according to Theorem 5, the types of quantum Fisher information fitting the Definition 3 are of the form Eq. (89). Both differ from the definitions of Refs. [15, 16] by a normalization factor. For which $f(x)$ is the quantum Fisher information $F_Q^f[\var, A]$ the largest, if it is normalized such that for pure states they equal to the usual variance, Eq. (1), and to four times the usual variance, respectively. Based on Eqs. (79), (80), (85) and (87), the two definitions are as follows.

$$\begin{align*}
\var_f(A) &= \frac{1}{2} \sum_{i,j} m_f(\lambda_i, \lambda_j) |A_{ij}|^2 - \sum |\lambda_i A_{ij}|^2, \\
F_Q^f[\var, A] &= 2 \sum_{i,j} m_f(1, 0) (\lambda_i - \lambda_j)^2 |A_{ij}|^2.
\end{align*}$$

(90)

For each mean $m_f(a, b)$, Equation [90] provides a generalized variance fulfilling Definition 1 and a generalized quantum Fisher information fulfilling Definition 3. It can be seen that the same $m_f(a, b)$ corresponds to the minimal variance and to the maximal quantum Fisher information. However, from Theorem 1, we already know that the minimal variance is the usual variance given in Eq. (1), corresponding to $f(x) = \frac{1+x}{2}$, i.e., to the arithmetic mean. Hence, $m_f(a, b) = \frac{a+b}{2}$. Hence, Theorem 3 follows.

**VI. CONCLUSIONS**

We have defined generalized variances and different types of quantum Fisher information. We have shown that the smallest variance is the usual variance since it is its own concave roof. We have also shown that the largest quantum Fisher information is the usual quantum Fisher information for rank-2 density matrices and for operators that have zero diagonal elements in the eigenbasis of the density matrix. In other words, the quantum Fisher information is 4 times the convex roof of the variance for such matrices. We found strong numerical evidence that the quantum Fisher information is very close to the
convex roof even for matrices with non-zero diagonal elements and for matrices with a larger rank. We related our findings to the generalized variances and types of quantum Fisher information defined in Refs. [15, 17]. We showed that the usual quantum Fisher information is the largest within the various types of generalized quantum Fisher information of Refs. [15, 16], if appropriate normalization is applied. Hence we conjectured that the usual quantum Fisher information is four times the convex roof of the variance for density matrices and Hermitian operators of any size.

In the future, it would be important to prove Conjecture 1. Since convex roofs appear in entanglement theory very often, our findings help to establish further connections between the quantum Fisher information and entanglement theory [16, 30, 33]. For instance, it would be important to examine the consequences of our findings concerning the detection of entanglement with variances [33] or the quantum Fisher information [6–9], or the estimations between the quantum Fisher information and entropies [33]. Since convex roofs appear in entanglement theory [16], if appropriate normalization is applied. Hence we conjectured that the quantum Fisher information of Refs. [15, 16], if appropriate normalization is applied. Hence we conjectured that the quantum Fisher information defined in Refs. [15, 17]. We showed that the usual quantum Fisher information is the largest within the various types of generalized quantum Fisher information of Refs. [15, 16], if appropriate normalization is applied. Hence we conjectured that the quantum Fisher information is four times the convex roof of the variance for density matrices and Hermitian operators of any size.

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Appendix A: Semidefinite programs used for obtaining the global optimum

Here we show how to compute $B_{SPPT}(\varrho, A) = B_{SE_{N}}(\varrho, A)$ and $B_{SE_{N}}(\varrho, A)$ defined in Eq. (69) using semidefinite programming.

Semi-definite programming can be used to look for the infimum over symmetric PPT states, i.e., finding $B_{SE_{N}}(\varrho, A)$ given in Eq. (69). In this way we obtain an upper bound on the infimum for separable states as has also been used in Ref. [31]. The corresponding task can be formulated as a standard semidefinite program as

\[
\begin{align*}
\text{minimize} & \quad 2\langle (A \otimes 1 - 1 \otimes A)^2 \rangle_{\varrho_0}, \\
\text{subject to} & \quad \varrho = \sum_{m,n} |s_m(2,d)\rangle \langle s_n(2,d)| \otimes \rho_{mn}, \\
& \quad \varrho_0 \geq 0, \quad \text{Tr}(\varrho_0) = 1, \\
& \quad \varrho_1^{T_1} \geq 0, \quad \text{Tr}_2(\varrho_0) = \varrho.
\end{align*}
\]

(A1)

Here $|s_n(2,d)\rangle$ denotes the basis states of the symmetric subspace for 2 qudits of dimension $d$. In order to ensure that the optimization is only over symmetric states, we define the density matrix $\varrho_s$, which is a two-qudit symmetric state given in the basis of the symmetric states. $\varrho_0$ is a density matrix given in the product basis, and $T_1$ means partial transpose according to the first subsystem.

Semi-definite programming can also be used to look for the infimum over symmetric states that have a PPT symmetric extension over $N$ qudits appearing in $B_{SE_{N}}(\varrho, A)$ defined in Eq. (69). The corresponding task can be formulated as a semidefinite program as

\[
\begin{align*}
\text{minimize} & \quad 2\langle (A \otimes 1 - 1 \otimes A)^2 \rangle_{\varrho_{12}}, \\
\text{subject to} & \quad \varrho_0 = \sum_{m,n} |s_m(N,d)\rangle \langle s_n(N,d)| \otimes \rho_{mn}, \\
& \quad \varrho_0 \geq 0, \quad \text{Tr}(\varrho_0) = 1, \\
& \quad \varrho_0^{T_1}_{12,\ldots, M} \geq 0 \text{ for } M = 1, 2, \ldots, \left\lceil \frac{N}{2} \right\rceil, \\
& \quad \varrho_{12} = \text{Tr}_{3,4,\ldots,N}(\varrho_0), \quad \text{Tr}_2(\varrho_{12}) = \varrho.
\end{align*}
\]

(A2)

where $\varrho_{12}$ is the two-qudit reduced state of $\varrho_0$ and $\lfloor x \rfloor$ is the integer part of $x$. Now $\varrho_0$ and $\varrho_{12}$ are $N$-qudit states. Here $|s_n(N,d)\rangle$ denotes the basis states of the symmetric subspace for $N$ qudits of dimension $d$. For $N$ qubits, these are

\[
|s_n(N, 2)\rangle = \binom{n}{N}^{-\frac{1}{2}} \sum_k P_k (|1\rangle^\otimes n |0\rangle^\otimes (N-n)),
\]

(A3)

for $n = 0, 1, \ldots, N$, where the summation is over all different permutations. Now $T_I$ means partial transposition according to the group of qudits given in $I$. Such an optimization has already been used in Ref. [32] for studying the entanglement properties of symmetric quantum states.

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