THE RADIO AND OPTICAL LUMINOSITY EVOLUTION OF QUASARS. II. THE SDSS SAMPLE

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ABSTRACT

We determine the radio and optical luminosity evolutions and the true distribution of the radio-loudness parameter $R$, defined as the ratio of the radio to optical luminosity, for a set of more than 5000 quasars combining Sloan Digital Sky Survey optical and faint images of the Radio Sky at Twenty cm (FIRST) radio data. We apply the method of Efron and Petrosian to access the intrinsic distribution parameters, taking into account the truncations and correlations inherent in the data. We find that the population exhibits strong positive evolution with redshift in both wavebands, with somewhat greater radio evolution than optical. With the luminosity evolutions accounted for, we determine the density evolutions and local radio and optical luminosity functions. The intrinsic distribution of the radio-loudness parameter $R$ is found to be quite different from the observed one and is smooth with no evidence of a bimodality in radio loudness for log $R \gtrsim -1$. The results we find are in general agreement with the previous analysis of Singal et al., which used POSS-I optical and FIRST radio data.

Key words: galaxies: active – galaxies: jets – methods: data analysis – methods: statistical – quasars: general

Online-only material: color figures

1. INTRODUCTION

Quasars are distant active galactic nuclei (AGNs) with emission seen across the electromagnetic spectrum, from radio to X-rays. The optical emission from quasars is thought to be dominated by the radiation from the accretion disk around supermassive black holes, while the radio emission is dominated by the plasma outflowing from the black hole/accretion disk systems. Because of this, different but complementary information can be gathered from both photon energy ranges regarding the cosmological evolution of AGNs. It is therefore important to determine in detail the redshift evolutions of quasars in both radio and optical regimes.

In a previous paper (Singal et al. 2011, hereafter QP1), we explored the luminosity evolutions and radio-loudness distribution of quasars with a data set consisting of the overlap of the FIRST (Faint Images of the Radio Sky at Twenty cm) radio survey with the Automatic Plate Measuring Facility catalog of the Palomar Observatory Sky Survey (POSS-I), as presented by White et al. (2000). In this paper, we present the results from a much larger data set consisting of the overlap of FIRST with the Sloan Digital Sky Survey (SDSS) Data Release 7 (Abazajian et al. 2009) quasar catalog, which contains a factor of $\sim 10$ more objects and spans redshifts from 0.065 to 5.46.

In the literature, the evolution of the quasar luminosity function (LF) has been described not only for optical and radio luminosities but also for X-ray, infrared, and bolometric luminosities (e.g., Ueda et al. 2003; Richards et al. 2006; Mateu et al. 2006; Hopkins et al. 2007; Croom et al. 2009). The shape of the LF and its evolution are usually obtained from a flux-limited sample in the $a$ waveband $f_a > f_{\text{lim},a}$ with $f_{\text{lim},a}$ denoting the flux limit and $L_a = 4\pi d_L^2f_a/K_a$, where $d_L(z)$ is the luminosity distance and $K_a(z)$ stands for the K-correction. For a pure power-law emission spectrum of index $\epsilon_a$ defined as $f_a \propto \nu^{-\epsilon_a}$, one has $K_a(z) = (1 + z)^{1-\epsilon_a}$. This simple form may be modified for optical data by the presence of emission lines, as in this work.

In general, the determination of the LF and its evolution requires analysis of the bi-variate distribution $\Psi_L(L_{\text{a}}, z)$. The first step of the process should be the determination of whether the variables of the distributions are correlated or are statistically independent. A correlation between $L_a$ and $z$ is a consequence of luminosity evolution. In the case of quasars with the optical and some other band luminosity, we have at least a tri-variate function. One must determine not only the correlations between the redshift and individual luminosities (i.e., the two luminosity evolutions) but also the possible intrinsic correlation between the two luminosities, before individual distributions can be determined. Knowledge of these correlations and distributions is essential not only for constraining robustly the cosmological evolution of active galaxies but also for interpretation of related observations, such as the extragalactic background radiation (e.g., Singal et al. 2010; Hopkins et al. 2010).

A related question is the distribution of the “radio-loudness parameter,” $R = L_{\text{rad}}/L_{\text{opt}}$, for the quasar population, defined as the ratio of the 5 GHz radio to 2500 Å optical luminosity spectral densities, and the distinction between so-called radio-loud ($R > 10$; RL for short) and radio-quiet ($R < 10$; RQ for short) quasars. Weak hints of a bimodality in the distribution of the radio-loudness parameter described by Kellerman et al. (1989) suggested that log $R = 1$ could be chosen as the radio-loud/radio-quiet demarcation value. Using this value for the division between RL and RQ quasars, the differences between the two classes have been investigated, including the possibility of discernible cosmological evolution of the RL and RQ populations (e.g., Miller et al. 1990; Goldschmidt et al. 1999; Jiang et al. 2007). Still, the more recent analyses of different samples of objects reported in the literature so far gave rather inconclusive results on whether any bimodality in the distribution of the radio-loudness parameter for quasars is inherent in the population (see...
Ivezic et al. 2002, 2004; Cirasuolo et al. 2003; Zamfir et al. 2008; Kimball et al. 2011; Mahony et al. 2012). In QP1 we found no evidence for a bimodality in $R$ in the range $-1 < \log R < 4$.

In addition to the above-cited works, there have been many papers dealing with this ratio and RL versus RQ issue, as well as luminosity ratios at other wavelengths: IR/radio, optical/X-ray, etc. However, to the best of our knowledge, except for QP1, none of these works have addressed the correlations between the radio and optical luminosities, which is necessary for such an analysis. Additionally, in general they have not concentrated on obtaining the intrinsic—as opposed to the raw observed—distribution of the ratio, which is related to the tri-variate LF $\Psi(L_{\text{opt}}, L_{\text{rad}}, z)$ by

$$G_R(R, z) = \int_0^\infty \Psi(L_{\text{opt}}, R, L_{\text{opt}}, z) \, L_{\text{opt}} \, dL_{\text{opt}} = \int_0^\infty \Psi \left( \frac{L_{\text{rad}}}{R}, L_{\text{rad}}, z \right) \, L_{\text{rad}} \, dL_{\text{rad}}. \quad (1)$$

In Appendix A of QP1 we showed how, even in the simplest cases, the observed distributions can be very different from the intrinsic ones. Thus, for determination of the true distributions the data truncations must be determined and the correlations between all variables must be properly evaluated.

Efron & Petrosian (1992, 1999, EP for short) developed new methods for determination of the existence of correlation or independence of variables from a flux-limited and more generally truncated data set, which were further expanded in QP1. Our aim in this paper is to take all the selection and correlation effects into account in determination of the true evolution of optical and radio luminosities and their ratio and to find their distributions, using the larger SDSS DR7 QSO × FIRST data set.

In Section 2 we describe the radio and optical data used. Section 3 contains a general discussion of luminosity evolution and the sequence of the analysis. In Section 4 we apply the EP method to achieve the luminosity–redshift evolutions and the correlation between the luminosities. We determine the density evolution in Section 5 and the local LFs in Section 6. A discussion of the radio-loudness distribution is presented in Section 7. In Section 8 we investigate some of the assumptions used, and Section 9 contains a discussion of the results. This work assumes the standard ΛCDM cosmology throughout, with $H_0 = 71$ km s$^{-1}$ Mpc$^{-1}$, $\Omega_\Lambda = 0.7$, and $\Omega_m = 0.3$.

2. DATA

In order to evaluate the luminosity evolution in both radio and optical and to separate and compare these effects, we require a data set that has both radio and optical fluxes to reasonable limits across a broad range of redshifts that contains a significant number of both RL and RQ objects. The overlap of the FIRST bright quasar radio survey (Becker et al. 1995; White et al. 1997) with the SDSS DR7 quasar catalog (Schneider et al. 2010) can form such a data set.

The SDSS DR7 quasar catalog has a limiting $i$-band magnitude of 21, or $f_{m,i} = 0.015$ mJy, and contains over 105,000 objects with redshifts ranging from 0.065 to 5.46 over 9380 deg$^2$. We work with a more restrictive $i$-band optical magnitude limit of 19.1 ($f_{m,i} = 0.083$ mJy) to form a parent optical catalog that has a smoother redshift distribution with a reduced bias toward objects with $z > 2$ and with a uniform limiting flux for every redshift (see Section 6 of DR7Q). The $i < 19.1$ catalog consists of 63,942 objects, with a maximum redshift of 4.98, shown in Figure 1.

In the SDSS DR7 catalog, objects are identified as quasar candidates either if they have the requisite optical colors or if they have a radio match within 2″. Only 1% of the quasars in the catalog were selected from the candidate list for spectroscopic follow-up based on the latter criterion. We also note the presence of an upper limit $i$-band magnitude of 15.0 for inclusion in the catalog. However, this criterion is not completely rigorous, as mentioned in DR7Q. Further properties of the DR7 quasar catalog, such as black hole masses obtained from line emission measurements, are presented in Shen et al. (2011).

The FIRST survey, carried out with the Very Large Array (VLA) in B configuration between 1995 and 2004, has a limiting peak pixel 1.4 GHz flux of 1 mJy and contains 816,000 sources over 9500 deg$^2$. Different criteria can be used to determine radio and optical matches to form a joint SDSS DR7 QSO × FIRST catalog. Jiang et al. (2007, hereafter J07), in their analysis of SDSS quasars, used a matching radius of 5″ for single radio matches and 30″ for multiple matches to an optical source. We construct a fiducial catalog using these radius matching criteria, but, unlike J07, we do not include quasars with no radio detection (see the discussion in Section 9.1).

This fiducial catalog contains 5677 objects ranging in redshift from 0.64 to 4.82 and with log $R$ ranging from −0.47 to 4.44. A total of 4327 of the objects in the fiducial catalog are single matches (which J07 calls “Fanaroff–Riley type I sources,” FR1s, perhaps incorrectly; see the discussion in Section 9.1), and 1350 are multiple matches (which J07 calls “Fanaroff–Riley type II objects,” or FR2s for short). This sample seems to be skewed.
toward RL as opposed to RQ quasars, with 4543 (80%) having $R > 10$. It should be noted that J07 used the SDSS Data Release 3, so our realization with the same matching criteria, even with the additional $i < 19.1$ magnitude cut and no objects without radio detections, results in more objects (J07 have 2566 objects).

We also construct a catalog with a universal 5" matching criterion, which we feel is more appropriate, as discussed in Section 9.1. In this catalog there are 5445 objects with only 25 multiple radio matches to an optical source. We will refer to this as the canonical sample, and base our analysis primarily on it, while comparing to results we obtain with the fiducial sample, which as discussed below are quite similar. In the canonical sample 4466 (82%) have $R > 10$. The major changes from the fiducial catalog to the canonical one are that some optical objects with multiple radio matches in the fiducial sample are reduced to having only single matches in the canonical sample, with a corresponding reduction in radio flux, while a small number of objects drop out entirely, having had multiple matches within 30" but none within 5". The fiducial sample and canonical sample are nearly identical in their redshift distribution, and the average radio luminosity differs only slightly, being 10% lower in the canonical sample.

As shown in Richards et al. (2006, hereafter R06; e.g., Figure 8 of that work), the SDSS optical quasar sample contains significant biases in the redshift distribution due to emission-line effects and the differing flux limits at $z > 2$. We have reduced the latter effect by using only the universally flux-limited $i < 19.1$ sample, and the former by adopting the full $\alpha_{\text{opt}} = 0.5$ power-law continuum plus emission-line $K$-corrections presented in R06 and discussed in Section 5 and Table 4 of that work. The methods of this work can then account for any bias resulting from emission-line effects, as long as they are included in the conversions from luminosity to flux, i.e., in the $K$-corrections. For the radio data we assume a power law with $\alpha_{\text{rad}} = 0.6$. Figures 2 and 3 show the radio and optical luminosities versus redshifts of the quasars in the canonical constructed SDSS × FIRST catalog. Figure 4 shows the radio-loudness parameter $R$ versus redshift for the canonical data set, while Figure 5 shows the radio luminosity versus the optical luminosity for different redshift bins.

3. GENERAL REMARKS ON LFs AND EVOLUTIONS

3.1. Luminosity and Density Evolution

The LF gives the number of objects per unit comoving volume $V$ per unit source luminosity, so that the number density is $dN/dV = \int dL \Psi_a(L_a, z)$ and the total number is $N = \int dL \int dz (dV/dz) \Psi_a(L_a, z)$. To examine luminosity evolution, without loss of generality, we can write an LF in some waveband $a$ as

$$\Psi_a(L_a, z) = \rho(z) \psi_a(L_a/g_a(z), \eta_a)/g_a(z),$$

where $g_a(z)$ and $\rho(z)$ describe the luminosity evolution and comoving density evolution with redshift, respectively, and $\eta_a$ stands for parameters that describe the shape (e.g., power-law

Figure 2. Same as Figure 1 but for the quasars in the canonical SDSS × FIRST data set used in this analysis (5445 objects). The black points are the RL objects, while the red points are the RQ objects. (A color version of this figure is available in the online journal.)

Figure 3. 1.4 GHz rest-frame absolute luminosity density for the quasars in the canonical SDSS × FIRST data set used in this analysis (5445 objects). To obtain the 1.4 GHz luminosity density, we use the luminosity distance obtained from the redshift and the standard cosmology and the standard $K$-correction. We assume a radio spectral index of 0.6. The black points are the RL objects, while the red points are the RQ objects. (A color version of this figure is available in the online journal.)

Figure 4. Redshift distribution of the ratio $R$ of rest-frame absolute luminosities at 5 GHz and 2500 Å for the quasars in the canonical SDSS × FIRST data set used in this analysis. The 5 GHz luminosity is obtained from the 1.4 GHz luminosity assuming a radio spectral index of 0.6. The black points are the RL objects, while the red points are the RQ objects. Also plotted are the observed fraction of objects with $R > 10$ (dotted curve) and $R > 100$ (solid curve). (A color version of this figure is available in the online journal.)
indices and break values) of the $a$-band LF. In what follows we assume a non-evolving shape for the LF (i.e., $\eta_a = \text{const}$, independent of $L$ and $z$), which is a good approximation for determining the global evolutions. Once the luminosity evolution $g_a(z)$ is determined using the EP method, we can obtain the mono-variate distributions of the independent variables $L_a' = L_a / g_a(z)$ and $z$, namely, the density evolution $\rho(z)$ and “local” LF $\psi_a$. The total number of observed objects is then

$$N_{\text{tot}} = \int_0^{z_{\text{max}}} dz \int_{L_{a}(z)}^{\infty} dL_a \rho(z) \frac{dV}{dz} \frac{\psi_a(L_a / g_a(z))}{g_a(z)}.$$

We consider this form of the LF for luminosities in different bands, allowing for separate (optical and radio) luminosity evolutions.

In the previous work (QP1) with the White et al. data set, we assumed a simple power law for the luminosity evolution

$$g_a(z) = (1 + z)^{\delta_4}.$$

However, with the SDSS data set, which contains objects past $z = 4$ and many more objects past $z = 3$, we need to use a more complicated parameterization:

$$g_a(z) = \frac{(1 + z)^{\delta_4}}{1 + (z / z_{\text{cr}})^{\delta_6}}.$$

With either definition of $g_a(z)$ that gives $g_a(0) = 1$, the luminosities $L_a'$ refer to the de-evolved values at $z = 0$, hence the name “local.”

We have verified the form of Equation (5) as a good fit to the evolution by considering the optical evolution in the large parent optical-only sample in narrow redshift bins with the appropriate simple form of Equation (4) and then verifying the form of the cumulative evolution over a larger range. We have determined the appropriate value for $z_{\text{cr}}$ and confirmed that the same exponent is appropriate in the numerator and denominator, using the large set of optical-only data and the methods of Section 4.1. With the exponents fixed to be the same, the optimal value of $z_{\text{cr}}$ is determined to be $3.7 \pm 0.3$, by considering the $1\sigma$ range of the best-fit evolution while letting that numerical factor vary. We would expect that if the truncations and correlations have been properly accounted for by the methods of this work, $g_{\text{opt}}(z)$ as determined from the simultaneous radio and optical data set should match that as determined from the parent optical-only data set, which is shown to be the case (see Section 4).

We discuss the determination of the evolution factors $g_a(z)$ with the EP method, which in this parameterization becomes a determination of $k_{\text{rad}}$, in Section 4. The density evolution function $\rho(z)$ is determined by the method shown in Section 5. Once these are determined, we construct the local (de-evolved) LF $\psi_a(L_a')$, shown in Section 6.

### 3.2. Joint LFs

In general, determination of the evolution of the LF of extragalactic sources for any wavelength band except optical involves a tri-variate distribution because redshift determination requires optical observations, which introduces additional observational selection bias and data truncation. Thus, unless redshifts are known for all sources in a radio survey, we need to determine the combined LF $\Psi(L_{\text{opt}}, L_{\text{rad}}, z)$ from a tri-variate distribution of $z$ and the fluxes in the optical and radio bands. If the optical and radio luminosities were statistically independent variables, then this LF would be separable, $\Psi(L_{\text{opt}}, L_{\text{rad}}, z) = \Psi_{\text{opt}}(L_{\text{opt}}, z) \times \Psi_{\text{rad}}(L_{\text{rad}}, z)$, and we would be dealing with two bi-variate distributions.

However, the luminosities in the different wavebands may be correlated. The degree and form of the correlation must be determined from the data, and as described below, the EP method allows us to determine whether any pair of variables are independent or correlated. Once it is determined that the luminosities are correlated (see Section 4.2), the question must be asked whether this luminosity correlation is intrinsic to the population or induced in the data by flux limits and/or similar luminosity evolutions with redshift. Determination of which is the case is quite intricate, as discussed in Appendix B of QP1, and has not been explored much in the literature. This will be the subject of a future work. Here we will simply consider both possibilities.

At one extreme, if the luminosity correlation is intrinsic and not induced, one should seek a coordinate transformation to define a new pair of variables that are independent. This requires a parametric form for the transformation. One can define a new luminosity that is a combination of the two, which we call a “correlation-reduced radio luminosity,” $L_{\text{corr}} = L_{\text{rad}} / F(L_{\text{opt}} / L_{\text{fid}})$, where the function $F$ describes the correlation between $L_{\text{rad}}$ and $L_{\text{opt}}$ and $L_{\text{fid}}$ is a fiducial luminosity taken here\(^6\) to be $10^{28}$ erg s$^{-1}$ Hz$^{-1}$. For the correlation

\(^6\) This is a convenient choice for $L_{\text{fid}}$ as it is lower than the lowest 2500 Å luminosity considered in our sample, but results do not depend on the particular choice of numerical value.
function we will assume a simple power law

\[ L_{\text{crr}} = \frac{L_{\text{rad}}}{(L_{\text{opt}}/L_{\text{fid}})^{\alpha}}. \]  

(6)

where \( \alpha \) is a bulk power-law correlation index to be determined from the data. This is essentially a coordinate rotation in the log–log luminosity space. As shown in Section 4 below, EP also prescribe a method to determine a best-fit value for the index \( \alpha \) that orthogonalizes the new luminosities. Given the correlation function, we can then transform the data (and their truncation) into the new independent pair of luminosities (\( L_{\text{opt}} \) and \( L_{\text{crr}} \)).

At the other extreme, if the correlation between the luminosities is entirely induced, then the LFs are separable as described above and the analysis can proceed from there. As mentioned, we will consider both possibilities here. It turns out that the results obtained in both cases are very similar, implying that the luminosity–redshift correlations in the analyzed sample are much stronger than any intrinsic luminosity–luminosity correlations (if present). Further mechanics of the procedure for intrinsically correlated luminosities is discussed in the Appendix.

In this work we are also interested in the distribution of radio-loudness parameter \( R \), specifically its de-evolved value \( R' = L_{\text{rad}}/L_{\text{opt}} \). In the case of uncorrelated radio and optical luminosities or an induced, non-intrinsic radio–optical luminosity correlation, the local distribution of \( R' \), denoted hereafter as \( G(R') \), can be constructed from the local optical and radio LFs as in Equation (1), with the redshift evolution function \( g_{\text{rad}} = \frac{g_{\text{rad}}}{g_{\text{opt}}} \).

In the case of an intrinsic radio–optical luminosity correlation, \( G_{\text{R}} \) and the evolution \( g_{\text{R}} \) can be constructed by Equations (A3) and (A4) given in the Appendix.

4. DETERMINATION OF BEST-FIT CORRELATIONS

We now describe results obtained from the use of the procedures described in Section 3 on the data described in Section 2. Here we first give a brief summary of the algebra involved in the EP method. We generally follow the procedures described in more detail in QP1. This method uses the Kendall tau test to determine the best-fit values of parameters describing the correlation functions using the test statistic

\[ \tau = \frac{\sum_j (R_j - \xi_j)}{\sqrt{\sum_j V_j}}. \]  

(8)

to test the independence of two variables in a data set, say, \((x_j, y_j)\) for \( j = 1, \ldots, n \). Here \( R_j \) is the dependent variable (y) rank of the data point \( j \) in a set associated with it. For untruncated data (i.e., data truncated parallel to the axes) the set associated with point \( j \) includes all of the points with a lower independent variable value \((x_k < x_j)\). If the data are truncated, one must form the associated set consisting only of those points of lower independent variable \( x \) value that would have been observed if they were at the \( x \) value of point \( j \) given the truncation. As an example, if we have one-sided truncations as in Figures 2 and 3, ignoring the upper optical flux limit for the moment, then the associated set \( A_j = \{ k : y_k > y_j, \ y_k < y_j \} \), where \( y_k \) is the limiting \( y \) value of data point \( j \) (see EP for a full discussion of this method).

If \((x_j, y_j)\) were independent, then the rank \( R_j \) should be distributed uniformly between 0 and 1 with the expectation value and variance \( \xi_j = (1/2)(n + 1) \) and \( V_j = (1/12)(n^2 + 1) \), respectively, where \( n \) is the number of objects in object \( j \)'s associated set. Independence is rejected at the \( m \sigma \) level if \( |\tau| > m \). To find the best-fit correlation, the \( y \) data are then adjusted by defining \( y'_j = y_j / F(x_j) \) and the rank test is repeated, with different values of parameters of the function \( F \).

4.1. Optical-only Data Set Luminosity–Redshift Correlation

With the optical-only data set (here we use the \( i < 19.1 \) set described in Section 2), determination of the luminosity–redshift correlation function \( g_{\text{opt}}(z) \) by Equation (5) reduces to determination of the value of the index \( k_{\text{opt}} \). As evident from Figure 1, the \( L_{\text{opt}} - z \) data are heavily truncated due to the flux limits of SDSS. The associated set for object \( j \) includes only those objects that are sufficiently luminous to exceed the optical flux minimum for inclusion in the survey if they were located at the redshift of the object in question, i.e., \( L_k \geq L_{\text{min}}(z_j) \). As the values of \( k_{\text{opt}} \) are adjusted and \( L_{\text{opt}} \) is scaled by \( g_{\text{opt}}(k, z) \), the luminosity cutoff limits for a given redshift are also scaled by \( g_{\text{opt}}(k, z) \).

Figure 6 shows the value of the test statistic \( \tau \) for a range of values of \( k_{\text{opt}} \) given the parameterization of Equation (5). It is seen that the best-fit value of \( k_{\text{opt}} \) is 3.3 with a \( 1 \sigma \) range of 0.1. This indicates that quasars have strong optical luminosity evolution, i.e., the LF shifts to higher luminosities at larger redshifts.

We have performed an optimization procedure to determine the best-fit value of \( z_{\text{cr}} \), the numerical constant in the denominator of Equation (5), as well as whether the same exponent is appropriate for the numerator and denominator. This was done by varying these parameters and determining the best-fit range of values in the same way as described above. We find \( z_{\text{cr}} = 3.7 \pm 0.3 \).

In this analysis we have ignored the upper optical flux limit of SDSS quasars corresponding roughly to an \( i \)-band magnitude of 15.0 as discussed in Section 2. The reason for this is that data truncations are only consequential in this analysis if the truncation is actually denying the sample data points that would otherwise be there. As can be seen in Figure 1, there are very few
objects approaching this upper truncation limit, indicating that the truncation is not consequential, i.e., it does not appreciably alter the sample from the underlying population.

4.2. Radio–Optical Luminosity Correlation

Turning to the combined radio and optical data set, we will determine the correlation between the radio and optical luminosities. The radio and optical luminosities are obtained from radio and optical fluxes from a two-flux-limited sample so that the data points in the two-dimensional flux space are truncated parallel to the axes that we consider to be untruncated. Since the two luminosities have essentially the same relationship with their respective fluxes, except for a minor difference in the \( K \)-correction terms, we can consider the luminosity data to also be untruncated in the \( L_{\text{opt}}-L_{\text{rad}} \) space. In that case the determination of the associated set is trivial and one is dealing with the standard Kendall tau test. Assuming the correlation function between the luminosities \( F(x) = x^\alpha \), we calculate the test statistic \( \tau \) as a function of \( \alpha \). Figure 7 shows the absolute value of the test statistic \( \tau \) versus \( \alpha \), from which we get the best-fit value of \( \alpha = 1.2 \) with a 1\( \sigma \) range \( \pm 0.1 \). The result is similar whether the canonical or fiducial combined data set is used. As expected, \( \alpha \) is near unity, and this result is compatible with that obtained in QP1 (\( \alpha = 1.3 \pm 0.2 \)).

As discussed in Section 3.2, this correlation may be inherent in the population or may be an artifact of the flux limits and wide range of redshifts. The general question of determining whether an observed luminosity correlation is intrinsic or induced will be explored in a future work. For this analysis going forward, we will consider both possibilities. We see that it does not make a significant difference for the major conclusions of this work.

4.3. Joint Data Set Luminosity–Redshift Correlations

The basic method for determining simultaneously the best-fit \( k_{\text{opt}} \) and \( k_{\text{rad}} \), given the evolution forms in Equation (5), is the same as described in Section 4.1, but in this case the procedure is more complicated because we now are dealing with a three-dimensional distribution \( (L_{\text{rad}}, L_{\text{opt}}, z) \) and two correlation functions \( (g_{\text{rad}}(z) \) and \( g_{\text{opt}}(z)) \).

Since we have two criteria for truncation, the associated set for each object includes only those objects that are sufficiently luminous in both bands to exceed both flux minima for inclusion in the survey if they were located at the redshift of the object in question. Consequently, we have a two-dimensional minimization problem, because both the optical and radio evolution factors, \( g_{\text{opt}}(z) \) and \( g_{\text{rad}}(z) \), come into play, as the luminosity cutoff limits for a given redshift are also adjusted by powers of \( k_{\text{opt}} \) and \( k_{\text{rad}} \).

We form a test statistic \( \tau_{\text{comb}} = \sqrt{\tau_{\text{opt}}^2 + \tau_{\text{rad}}^2} \), where \( \tau_{\text{opt}} \) and \( \tau_{\text{rad}} \) are those evaluated considering the objects’ optical and radio luminosities, respectively. The favored values of \( k_{\text{opt}} \) and \( k_{\text{rad}} \) are those that simultaneously give the lowest \( \tau_{\text{comb}} \), and, again, we take the 1\( \sigma \) limits as those in which \( \tau_{\text{comb}} < 1 \). For visualization, Figure 8 shows a surface plot of \( \tau_{\text{comb}} \). Figure 9 shows the best-fit values of \( k_{\text{opt}} \) and \( k_{\text{rad}} \) taking the 1 and 2\( \sigma \) contours, for the canonical radio–optical data set. We have verified this method with a simulated Monte Carlo data set as discussed in QP1.
We see that positive evolution in both radio and optical wavebands is favored. The minimum value of $\tau_{\text{comb}}$ favors an optical evolution of $k_{\text{opt}} = 3.5$ and a radio evolution of $k_{\text{rad}} = 5.5$, with a very small uncertainty at the 1$\sigma$ level. The fact that $k_{\text{opt}}$ as determined here from the combined data set is equal to that determined from the much larger optical-only data set (see Section 4.1 and Figure 6) indicates that the truncations have been properly handled.

In the above analysis we have assumed sharp truncation boundaries and that the data are complete above the boundaries. As discussed in Section 8, this may not be the case for the FIRST radio data. If we restrict the sample to only data with $f_{\text{rad}} > 2$ mJy, the favored optical evolution range for the canonical data set lowers slightly to $k_{\text{opt}} = 3.0 \pm 0.5$ and the favored radio evolution range lowers slightly to $k_{\text{rad}} = 5.0 \pm 0.25$. These overlap with the results considering the entire combined canonical sample at the 1$\sigma$ level.

These results indicate that quasars have undergone a significantly greater radio evolution relative to optical evolution, in general agreement with the result in QP1, although the best-fit differential evolution was even more dramatic there, given the form of $g(z)$ used.

The results for the best-fit $k_{\text{opt}}$ and $k_{\text{rad}}$ are identical if the different fiducial radio–optical data set matching radius criteria are used, indicating that the addition of extra radio flux and a few additional sources do not matter for this determination. If we consider that the radio–optical luminosity correlation is inherent in the data, then the orthogonal luminosities are $L_{\text{opt}}$ and $L_{\text{rad}}$ (see Section 3.2) and we can determine the best-fit evolutions $g_{\text{opt}}(z)$ and $g_{\text{rad}}(z)$. These results favor $k_{\text{opt}} = 3.5$ and $k_{\text{rad}} = 2$. In this case the best-fit radio evolution can be recovered by Equation (A2) in the Appendix, and it is $k_{\text{rad}} = 5.5$, in perfect agreement with the results obtained from considering $k_{\text{opt}}$ and $k_{\text{rad}}$ as orthogonal, indicating the robustness of the result.

5. DENSITY EVOLUTION

Next we determine the density evolution $\rho(z)$. One can define the cumulative density function

$$\sigma(z) = \int_{0}^{z} \frac{dV}{dz} \rho(z) dz,$$

which, following Petrosian (1992) based on Lynden-Bell (1971), can be calculated by

$$\sigma(z) = \prod_{j} \left(1 + \frac{1}{m(j)} \right),$$

where $j$ runs over all objects with a redshift lower than or equal to $z$, and $m(j)$ is the number of objects with a redshift lower than the redshift of object $j$ that are in object $j$’s associated set. In this case, the associated set is again those objects with sufficient optical and radio luminosity that they would be seen if they were at object $j$’s redshift. The use of only the associated set for each object removes the biases introduced by the data truncation. Then the density evolution $\rho(z)$ is

$$\frac{d\sigma(z)}{dz} = \frac{1}{dV/dz}.$$ 

However, to determine the density evolution, the previously determined luminosity evolution must be taken out. Thus, the objects’ optical and radio luminosities, as well as the optical and radio luminosity limits for inclusion in the associated set for given redshifts in the calculation of $\sigma$ by Equation (10), are scaled by taking out factors of $g_{\text{opt}}(z)$ and $g_{\text{rad}}(z)$, determined as above. Figures 10 and 11 show the cumulative and differential density evolutions, respectively. The number density of quasars seems to peak at between redshifts 1 and 1.5, earlier than generally thought for the most luminous quasars (e.g., Shaver et al. 1996), and earlier than that found in R06, but more similar to the peak found for less luminous quasars by Hopkins et al. (2007), and in agreement with Maloney & Petrosian (1999). This is slightly earlier than we found in QP1.

The normalization of $\rho(z)$ is determined by Equation (3), with the customary choice of $\int_{0}^{\infty} \psi(L') dL' = 1$. As stated in R06, the main sources of bias in the redshift distribution of the SDSS quasar sample are (1) the differing magnitude limits for $z > 2$, (2) the effects of emission lines on $i'$-band flux at different redshifts, and, at a somewhat less important level, (3) the inclusion of extended sources at the lowest redshifts.

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We note a notional difference with QP1 where we plotted $\rho(z) = d\sigma/dz$ rather than $\rho$ as defined in Equation (11) here with a factor of $dV/dz$ taken out.

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equal to $La_{R06}$, dealt with issue 1 by restricting the sample to a universal $\rho$ or deviation in $K$-evolution luminosity function $\psi_a$. The higher than the luminosity of object $k$ as a function of redshift, by the prime indicates that the luminosity evolution has been taken provided in Table 1.

Following Petrosian (1992), we can form the luminosity density functions in Figure 12. As expected, because of the greater luminosity evolution in the radio band, the radio luminosity density peaks earlier in redshift than the optical luminosity density.

### 6. LOCAL LFs

#### 6.1. General Considerations

In a parallel procedure we can use the local (redshift evolution taken out, or “de-evolved”) luminosity distributions (and de-evolved luminosity thresholds) to determine the “local” LFs $\psi_a(L'_a)$, where again the subscript $a$ denotes the waveband and the prime indicates that the luminosity evolution has been taken out. We first obtain the cumulative LF

$$\Phi_a(L'_a) = \int_{L'_a}^{\infty} \psi_a(L'_a) dL'_a. \quad (13)$$

Following Petrosian (1992), $\Phi_a(L'_a)$ can then be calculated by

$$\Phi_a(L'_a) = \prod_k \left(1 + \frac{1}{n(k)} \right), \quad (14)$$

where $k$ runs over all objects with a luminosity greater than or equal to $L'_a$, and $n(k)$ is the number of objects with a luminosity higher than the luminosity of object $k$ that are in object $k$’s associated set, determined in the same manner as above. The luminosity function $\psi_a(L'_a)$ is

![Figure 12. Optical (solid curve) and radio (dashed curve) luminosity densities $\Phi_{opt}(z)$ and $\Phi_{rad}(z)$, discussed in Section 5 and Equation (12). Note the different scales.](image)

#### 6.2. Local Optical Luminosity Function

Figures 13 and 14 show the local cumulative $\Phi_{opt}(L'_opt)$ and differential $\psi_{opt}(L'_opt)$ local optical LFs of the quasars in the canonical combined data set and that of the optical-only data set.

The optical LF shows possible evidence of a break at $\sim 10^{30}$ erg s$^{-1}$ Hz$^{-1}$, which was present already in data analyzed.
in Petrosian (1973) and was also seen in QP1. With the optical-only data set, fitting a broken power law yields values $-2.8 \pm 0.2$ and $-4.1 \pm 0.4$ below and above the break, respectively. With the canonical combined radio–optical data set, fitting a broken power law yields values $-2.8 \pm 0.3$ and $-3.8 \pm 0.5$. If we allow for the possibility of additional uncertainty resulting from the consideration of possible radio incompleteness at faint fluxes (see discussion in Section 8), the range on the power law below the break changes to $-2.1 \pm 0.1$, and above the break the results are not affected.

As the optical LF has been studied extensively in various AGN surveys, we can compare the slope of $\psi_{\text{opt}}(L_{\text{opt}})$ obtained here to values reported in the literature. In QP1, we found power-law values of $-2.0 \pm 0.2$ and $-3.2 \pm 0.2$ below and above the break, respectively. Boyle et al. (2000), using the 2dF optical data set (but with no radio overlap criteria), use a customary broken power-law form for the LF, with values ranging from $-1.39$ to $-3.95$ for different realizations, showing reasonable agreement.

### 6.3. Local Radio Luminosity Function

Figure 15 shows the local radio LF $\psi_{\text{rad}}(L_{\text{rad}})$. It is seen that the local radio LF contains a possible break around $2 \times 10^{31}$ erg s$^{-1}$ Hz$^{-1}$, with a power-law slope of $-1.5 \pm 0.1$ below the break and $-2.6 \pm 0.1$ above the break. The range for the power law above the break is increased slightly to $-2.5 \pm 0.3$ if the effects of possible radio incompleteness are included, as in Section 8, while below the break it is unchanged, indicating that this is a small effect. It is interesting to note that the value of the radio-loudness parameter $R$ corresponding to the break in the radio and optical LFs is very close to the critical value $R \sim 10$ widely discussed in the context of the RL/RQ dichotomy.

We can also construct the local radio LF from $\psi_{\text{opt}}(L_{\text{opt}})$ and $\psi_{\text{opt}}(L_{\text{rad}})$ with Equation (A1) of the Appendix, under the assumption that the radio and optical luminosity correlation is intrinsic. This is shown by the small black points in Figure 15.

It is seen that $\psi_{\text{rad}}(L_{\text{rad}})$ constructed in this way is nearly identical to that determined by considering the radio and optical luminosity correlation to be induced, indicating that the radio luminosity function result is robust to this consideration.

In QP1 for the local radio luminosity function we found a break at $10^{31}$ erg s$^{-1}$ Hz$^{-1}$, with a power-law slope of $-1.7 \pm 0.1$ below the break and $-2.4 \pm 0.1$ above it, in excellent agreement with the results here. The slope above the break seen here is similar to earlier results of Schmidt (1972) and Petrosian (1973), which probed only high luminosities. A more complete comparison can be made with Mauch & Sadler (2007), who form radio LFs of local sources in the Second Incremental Data Release of the 6 degree Field Galaxy Survey (6dFGS) radio catalog. For the sources they identify as AGNs, they find a break at $3.1 \times 10^{31}$ erg s$^{-1}$ Hz$^{-1}$, with slopes of $-2.27 \pm 0.18$ and $-1.49 \pm 0.04$ above and below the break, respectively, in good agreement with our results.

In Figure 15 we also show the radio LFs for the same range of $L_{\text{rad}}$, as determined by Kimball et al. (2011) for two samples; one for a sample of SDSS optical and NRAO Very Large Array Sky Survey (NVSS) radio data, and another for deep radio observations with the Extended Very Large Array in which almost every SDSS quasar in a narrow redshift range in a field was detected in radio. The results presented there are the luminosity function only for sources in the redshift range $0.2 < z < 0.3$, which we have scaled by means of Equation (5) and converted from 6 GHz to 1.4 GHz with a radio spectral index of 0.6 to obtain the local luminosity function at 1.4 GHz for direct comparison with our results. It is seen that our LFs agree on the faint end of the range considered here but potentially disagree only on the very bright end.
As discussed in Section 3, we can reconstruct the local distribution of $R$ to be weighted toward much higher values of $R$ given Equation (A4). The change in the distribution is significant and systematically biased. The naive, raw distribution instead of the reconstructed intrinsic one increasing redshift is also shown in Figure 16.

Another way to look at this is that we have found that the radio luminosity would likely be wrong. In Figure 16 we show this raw distribution of $R$ by the triangles, arrived at by using the raw values of $R$ from the data and forming a distribution in the manner of Equations (14) and (15) with no data truncations accounted for. As discussed in Section 3, we can reconstruct the local distribution of $G_R(R')$, as in Equation (1), which provides for a more proper accounting of the biases and truncations and is an estimate of the true, intrinsic distribution. The results of this calculation are also shown in Figure 16 by the filled circles. We include a range of error for the extremal possibilities that the correlation between the radio and optical luminosities is entirely independent of redshift, the raw observed distribution of $R$ toward increasing redshift can be validated in a simple way by examining the median value of $R$ versus redshift in bins of redshift, which indeed shows a steady increase.

This differential evolution is in disagreement with the result presented by J07, who show a decrease in fraction of RL sources with increasing redshift, which could be the case if the radio luminosity were to evolve more slowly than the optical luminosity. They, however, do not determine individual evolutions or LFs. On the other hand, Miller et al. (1990) have noted that the fraction of RL quasars may increase with redshift, which they attribute to a difference in the evolutions of the two populations (RL and RQ). Donoso et al. (2009) compute radio and optical LFs at different redshifts and reach the same conclusion. Cirasuolo et al. (2006) also find that the radio-loud fraction may modestly increase at high redshift. Although not directly comparable, LaFranca et al. (2010) show a similar evolution for $R_L$, the ratio of radio to X-ray luminosity, as we show here for $R$. As discussed in Section 9.1, we believe that our differences with JO7 stem from our inclusion of only radio-detected quasars in the joint radio–optical sample.

We note that our results favor one population in the range of $R$ considered here, in the sense that the distribution of $G(R)$, recovered from considering the data truncations inherent in the survey and correlations between the luminosities, is continuous and without any bimodality. This is in agreement with the conclusion we reached in QP1. One may hypothesize about a bimodality that is not centered on the commonly assumed value of $R = 10$ or above that value, but at a lower value of $R$ below those values probed in this analysis. Our results disfavor that as well. As seen in Figures 14 and 15, the lack of a significant change in the shapes of the computed local optical or radio LFs when the radio flux limits for inclusion in the analysis are changed, or the parent optical-only data set is considered, argues against the presence of an additional population of quasars that would become more prominent below the lowest values of $R$ considered here. The results of Kimball et al. (2011), which probe even lower radio fluxes and extend the radio luminosity function to lower values, do not show a large enough population of low-$R$ quasars that would produce a bimodality in the distribution of $R$.

In Figure 17 we plot our results along with two determinations of the radio observed $R$ distribution available in the literature, that of Ivezić et al. (2004) and that of Cirasuolo et al. (2006). These determinations from the literature are scaled to the ratio of 5 GHz radio to 2500 Å optical luminosity assuming optical and radio spectral indices of 0.5 and 0.6, respectively. In the former case (red and orange squares in the figure), the data set is from an SDSS × FIRST sample that is flux limited in radio and optical as in this work. The raw $R$ distribution seen there is similar to the one shown here, although with a more pronounced dip at $R \sim 10$. In the latter case (blue stars in the figure), the data are from a small patch of sky with deep VLA observations where the sample of SDSS quasars has 100% completeness down to $S_{1.4 \text{ GHz}} = 60 \mu$Jy, and with such radio completeness we would expect the observed distribution to be consistent with our reconstructed intrinsic distribution at low values of $R$, which evolves at a different rate than the optical luminosity, with the consequence that their ratio is a function of redshift. The radio loudness of the population increases by a factor of 3 by redshift 1, a factor of 8 by redshift 3, and a factor of 11 by redshift 5. This is a less dramatic evolution of the ratio than we found in QP1, where we used a simpler parameterization (see Section 3.1), although still quite significant. The general trend toward increasing $R$ with redshift can be validated in a simple way by examining the median value of $R$ versus redshift in bins of redshift, which indeed shows a steady increase.

### Figure 16.

Local distribution $G_R(R)$ in the 5 GHz radio to 2500 Å optical luminosity ratio $R$ for the quasars in the data set, plotted as $R \times G_R(R)$. The curves are from $G_R(R')$ as determined by the method of Equation (A3), taking account of the truncations and correlations in the luminosity evolutions, while the triangles result from forming a distribution with a naive use of the objects’ raw ratio. The error bars represent the extremal possibilities of the radio–optical luminosity correlation being entirely induced (open circles) and entirely intrinsic (closed circles). The normalization is arbitrary. Also shown is the proper radio-loudness distribution $G_L(R, z)$ at redshifts $z = 1$ (dashed line), $z = 3$ (dash-dotted line), and $z = 5$ (dotted line) evolved according to the form of Equation (A4). A comparison with data from the literature is shown in Figure 17.

7. DISTRIBUTION OF RADIO-LOUDNESS RATIOS

As stated in the introduction, naively one may expect that the ratio $R$ is independent of cosmological model and nearly independent of redshift, the raw observed distribution of $R$ would provide a good representation of the true distribution of this ratio. In Appendix A of QP1 we show that this assumption would likely be wrong. In Figure 16 we show this raw distribution of $R$ by the triangles, arrived at by using the raw values of $R$ from the data and forming a distribution in the manner of Equations (14) and (15) with no data truncations accounted for. As discussed in Section 3, we can reconstruct the local distribution of $G_R(R')$, as in Equation (1), which provides for a more proper accounting of the biases and truncations and is an estimate of the true, intrinsic distribution. The results of this calculation are also shown in Figure 16 by the filled circles. We include a range of error for the extremal possibilities that the correlation between the radio and optical luminosities is entirely induced or entirely intrinsic (open circles). In the latter case, $G_R(R')$ is given by Equation (A3) of the Appendix.

The distribution $G_R(R')$ calculated from the data taking into account the evolutions and truncations is clearly different from the raw distribution and shows no evidence of bimodality in the range of $R$ considered. The raw distribution is seen to be weighted toward much higher values of $R$. Using the naive, raw distribution instead of the reconstructed intrinsic one would result in estimates of quasar properties that would be significantly and systematically biased.

We also know the best-fit redshift evolution of the ratio, given Equation (A4). The change in the distribution of $R$ with increasing redshift is also shown in Figure 16.9 Another way to look at this is that we have found that the radio luminosity

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9 Note that we have not included the density evolution, which will shift the curves vertically but not change their shape.
indeed it is, but not necessarily at high values of $R$, where there may be sources missing, either due to cosmic variance or because they have significant radio flux but insufficient optical flux to be included in the SDSS sample (see the discussion in Section 9).

8. TESTS OF ASSUMPTIONS

Luminosity-dependent density evolution. One may be concerned that luminosity-dependent density evolution (LDDE), which is not considered in the functional forms for the LF used here, may be necessary to represent the evolution of the LF. As a test of this scenario, we divide the data into high and low halves of de-evolved luminosity $L^*$ (cutting on optical luminosity) and check the similarity of the computed density evolutions for the two sets. The density evolutions computed for both cuts are similar to each other, with the high-redshift half peaking in $\rho(z)$ at $z = 1.3$ and the low-redshift half peaking in $\rho(z)$ at $z = 1.1$. Given the similarity of these distributions to each other and to that computed from the data set as a whole, we conclude that we are justified in neglecting LDDE for the purposes of this analysis.

Optical measurement errors. There is the possibility that errors in the optical magnitudes could lead to a bias that could affect the results (Eddington 1940). The bias introduced by these errors would be negligible if the source counts as a function of flux, i.e., the log $N$−log $S$ distribution, were flat (i.e., $N(>S) \propto S^0$ and $dN/dS \propto S^{-1}$). Since the number density of sources increases with decreasing flux, it is more likely that a source will be included than excluded. However, the magnitude of this effect will depend on the faint-end source counts slope, and the shallower the slope, the smaller the effect. For constant fractional flux measurement errors, an error will be introduced on the normalization of the source counts, and therefore the LFs, and can be approximated by $1/2 \delta \times m_{\text{below}} (m_{\text{below}} + 1)$ (Teerikorpi 2004), where $\delta$ is the fractional error in flux and $m_{\text{below}}$ is the faint-end differential source counts power-law slope. To be conservative, we will adopt errors of 0.2 mag in $i$-band, to account for the intrinsic rms scatter due to source variability, although the typical reported SDSS DR7Q measurement errors are lower, on the order of a few hundredths of a magnitude. For the faint-end magnitudes of 19.1 and this error, $\delta$ is around 0.16. For $m_{\text{below}} \sim 2$, the bias on the luminosity function normalization will be around 7.6%. On the other hand, there will be an effect on the reconstructed slope of the luminosity function only if the fractional measurement errors change with luminosity. The data show only a modest dependence on $i$-band reported error with magnitude at magnitudes below 19.1, with magnitude errors at most a factor of 1.5 higher at the highest magnitude end of that range than for the lowest magnitudes, which corresponds to a small fractional flux error. Because data at a wide range of luminosities correspond to a given flux, we consider this effect to be negligible, even if scatter due to variability has some flux dependence.

Radio incompleteness. Lastly, the selection function for the FIRST objects used here might not be a sharp Heavyside function at a peak pixel flux density of 1 mJy, but rather smeared out. According to Figure 1 of J07, the selection function of FIRST for SDSS optically identified quasars in SDSS data release 3 is such that at an integrated flux density of 1 mJy only about 55% of sources are seen, and this number rises to 75% at 1.5 mJy and about 85% at 2 mJy. We have considered the sample to be limited by the peak pixel flux (i.e., surface brightness limited) in the radio rather than being limited by the integrated flux, in accordance with the criteria set forward in Becker et al. (1995).

The way to test the effects on our analysis of possible radio incompleteness is to repeat the analysis limiting the sample to a higher radio flux, where the sample would presumably be more complete, and determine the extent to which the calculated parameters change in a systematic way. We have done so with a lower radio flux limit of 2 mJy (4251 objects in the canonical sample), as opposed to the original 1 mJy (5445 objects in the canonical sample). The effect, propagated through the analysis, is quite minor. The 1σ uncertainties on $k_{\text{opt}}$ and $k_{\text{rad}}$ are extended slightly. There is no discernible effect on the correlation parameter $\alpha$. There is a negligible effect on the density evolutions, a small effect on the local LFs that are reported in Section 6, and a small effect on the shape of the $G_R(R)$ distribution that is sub-dominant to the effect of considering the radio–optical luminosity correlation to be intrinsic or induced. To the extent that faint flux radio incompleteness is present in the sample considered here, it does not seem to have a large systematic effect on the determination of the parameters in, and conclusions of, this analysis.

9. DISCUSSION

We have used a general and robust method to determine the radio and optical luminosity evolutions simultaneously for the quasars in the SDSS DR7 QSO × FIRST data set, which combines 1.4 GHz radio and $i$-band optical data for thousands of quasars ranging in redshift from 0.64 to 4.82 and over five orders of magnitude in radio loudness.

As can be seen, the results for the SDSS DR7 QSO × FIRST data set employed here are similar in bulk to the results for the White et al. data set that we presented in QP1. In this

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10 The fuzzyness of the truncation boundary has a similar effect as the data measurement errors in the sense that it is unimportant for $m_{\text{below}} = -1$ and more important for larger deviations from this value.
work we show that, importantly, the major conclusions are not sensitive to whether the correlation between the radio and optical luminosities is considered to be induced by similar redshift evolution or intrinsic. In the previous work, we assumed the correlation to be intrinsic. Additionally, the luminosity evolution found for the optical-only data set matches that found for the combined radio–optical data set, providing a test that the technique has properly handled the truncations and correlations inherent in the data.

We have seen that the results of this analysis are nearly identical for the two different radio–optical matching radius criteria considered: (1) that adopted by J07 with a 5″ radius for single matches and a 30″ radius for multiple matches, and (2) a universal 5″ matching radius, which we feel is more appropriate. The reason for this preference is that 30″ radius corresponds to a physical scale >100 kpc projected, at any redshift z > 0.2. Extended lobes in radio galaxies are known to reach similar or even larger sizes. These are, however, objects viewed at large inclinations, unlike quasars observed, by definition, at much smaller viewing angles (Barthel 1989). With such small viewing angles, projection effects are expected to limit the projected sizes of the extended structures in radio-loud quasars. In addition, evolutionary effects may play a role as well, with distant quasars being typically younger, and therefore more compact in radio, than nearby evolved radio galaxies. This is supported by a detailed investigation of radio morphologies of SDSS (DR3) × FIRST quasars by De Vries et al. (2006), who noted that an overwhelming majority of the selected sources are characterized by compact radio morphologies. For all these reasons, we believe that applying a universal 5″ matching radius criterion in assigning FIRST counterparts to the SDSS bona fide quasar sources is more appropriate, while the 30″ criterion for multiple matches has a high probability of including radio flux from objects unrelated to the optical quasar that they are supposedly linked to.

9.1. Differential Evolutions toward Radio Loudness with Increasing Redshift

Quasars are seen to exhibit positive luminosity evolution in both wavebands, with stronger positive evolution in radio than optical. The differential evolution can be seen in the evolution of $G_{R}(R')$ as shown in Figure 16. This is in agreement with the results presented for a combined FIRST × POSS-I data set in QP1, but in likely disagreement with J07, which claimed that the fraction of RL quasars decreases with increasing redshift. However, Miller et al. (1990) have noted that the fraction of RL quasars may increase with redshift, and Donoso et al. (2009) reached the same conclusion by computing radio and optical LFs at different redshifts. Cirasuolo et al. (2006) also found that the RL fraction of quasars may modestly increase at high redshift, and the conclusions of Balokovic et al. (2012) favor $R$ increasing with redshift. We note that a trend toward increasing radio loudness with redshift is visible to the eye even in the raw observed data as in Figure 5, although these data suffer from biases so that the true evolution of $R$ can only be recovered with an analysis method such as employed here.

We attribute our apparent disagreement with the J07 result to three causes: (1) they present one particular moment of the $R$ distribution—the “radio-loud fraction”—in bins, while we are presenting the distribution itself, so direct comparison is rather difficult; (2) they are calculating this fraction by including optical sources that have no radio detection, which is a potentially correctable bias; and (3) they are not including the effect of neglecting sources that are radio bright enough to appear in FIRST but do not appear in the SDSS quasar sample that they use.

There are potentially many AGN sources in the FIRST catalog that are not present in the SDSS quasar sample, due to their low optical fluxes. The sample used in J07 to calculate the radio-loud fraction in bins would be missing this population of sources with detectable radio emission but lower optical fluxes than would allow an SDSS detection or classification as a quasar source, while we have used an analysis that takes both flux limits properly into account. While for the joint data set the present work has used only quasars detected in both radio and optical and including both optical-color-selected and radio-match-selected sources, J07 used as their sample quasars detected in optical regardless of a radio detection, and their quasars were selected based solely on optical color.

Stepping back, there are four basic options when constructing a combined optical-radio data set of quasars for an analysis of the radio-loudness distribution and its evolution, given that many more confirmed quasars have an optical detection than a radio one:

I. quasars detected in both radio and optical
II. quasars detected in optical regardless of a radio detection
A. optical quasar candidates selected for follow-up based on either optical colors or presence of a radio match
B. optical quasar candidates selected for follow-up based only on optical colors.

Under the above rubric, we have used IA for the joint radio–optical data set, while J07 used IIB.

We believe that while using option II is appropriate to investigate the optical luminosity function and its evolution, such a data set is not appropriate for investigating the radio-loudness distribution and its evolution. This is because with such a data set one will always be introducing a bias by not including objects bright in radio but dim in optical, and at the same time not being able to say if such a population exists at all by means of some censored data method. The situation could be different if one were able to select two quasar samples based separately on the radio and optical data, down to given radio and optical flux limits of the surveys, and then to match the two samples forming a master-list with both radio and optical fluxes provided for each object included. Only then could non-detections in either radio or optical bands be considered and the censored data method be applied (see in this context, e.g., Feigelson & Berg 1983; Feigelson & Nelson 1985; Isobe et al. 1986). However, since an optical identification is needed to claim a quasar nature of a source in the first place, this is not possible, so in using an only optically detected set one will be always biased against a radio-loud population. On the other hand, if, as in the present work, one uses option I and then analyzes the data in such a way as to take the flux limits into account, an unbiased reconstruction of the intrinsic properties of the population can be achieved.

The question of whether option A or B is most appropriate is less straightforward. However, quasar optical color may be correlated with radio loudness (Kimball et al. 2011). If this is the case, then by selecting quasars based on the optical colors solely, one may introduce some bias when dealing with the radio-loudness distribution and evolution. Option IA means that the color criterion is not important: by including only those quasars that are detected in radio, and by allowing for radio-selected optical quasar candidates, we basically end up with
a radio-selected (and not color-selected) and radio and optical flux-limited sample of quasars.

9.2. Lack of Bimodality in Radio Loudness

Our results favor a single population of quasars with no bimodality in the radio-loudness parameter for the range of R considered here, in agreement with the conclusions of QP1. Some physical implications of a single population in this range of R are discussed in that work, including that there is thus no evidence for the existence of two physically different populations of quasars with respect to the radio/jet production efficiency. Although it is a somewhat different issue, there have also been studies of radio populations investigating the distribution of quasars, Seyferts, and other types of AGNs in the plane of radio loudness versus accretion rate (note that by definition all the quasars accrete at high and very high rates, unlike radio galaxies or Seyferts). These results indicate that there may be two “tracks” in the space, characterized by a similar non-monotonic dependence of the radio-loudness parameter on the Eddington ratio, but differing in normalization, and corresponding to what might be considered as an RL population hosted exclusively by elliptical galaxies, and an RQ population of AGNs hosted by both types of galaxies (Sikora et al. 2007). Such studies were, however, done—and in fact had to be done—on incomplete and inhomogeneous samples of AGNs and were claimed to depend on whether core or total radio luminosities are used (e.g., Broderick & Fender 2011). The host-related bimodal distribution of radio loudness in the entire AGN population is, however, distinct from the main problem addressed in this work, dealing strictly with a well-defined population of quasar-type AGNs (for which, notably, we use the FIRST—i.e., the total—radio fluxes). And in particular, the conclusion that there is no radio bimodality for the quasar population does not contradict the findings presented in Sikora et al. (2007), as these authors emphasized that while a complete sample of elliptical-hosted AGNs (including quasars) is expected to show continuous distribution of the radio loudness down to the “radio-quiet” regime, the point is that the spiral-hosted AGNs do never reach high values of the radio-loudness parameters (i.e., values comparable to those of radio galaxies), even though at very low accretion rates they are often characterized by \log R \gg 1 (see Section 3 of that work and the discussion therein).

In Section 7 we explored additional arguments against the presence of a bimodality in R at lower values of R, below those considered in this analysis. The caution to the above conclusion is that the sample analyzed here is dominated by RL objects, as we do not, for the reasons discussed above, consider SDSS quasars with no detected FIRST counterparts. Figure 5 indicates that with this SDSS × FIRST data set we are sensitive to regions where additional RQ objects would lie were they to exist (e.g., below the R = 10 line and above the \( L_{\text{rad,min}} \) lines in that figure), but there does not appear to be an overabundance of those objects. However, we cannot formally exclude the possibility that there may be a population of very radio-quiet quasars that would form a separate branch in Figure 16 were the abscissa extended to the left, resulting in a bimodal distribution of radio loudness for the entire quasar population.

Such a numerous population of very radio-quiet quasars, even if absent in the FIRST database, would not be missed in deeper radio surveys probing sub-mJy flux levels. In fact, Kimball et al. (2011) have used EVLA observations to detect 179 SDSS DR7Q quasars in the range 0.2 \( \lesssim z \lesssim 3 \) in their field down to a 6 GHz flux limit of 20 \( \mu \text{Jy} \). They claim to detect almost every quasar (97%) in the field in radio and construct the 6 GHz radio luminosity function with the \( V / V_{\text{max}} \) method. Their radio luminosity function is well fit with two populations: one with radio emission dominated by AGN central engine emission, and one with radio dominated by star formation. The latter becomes prominent at luminosities below \( L_{\text{6GHz}} \sim 10^{30} \text{ erg s}^{-1} \text{ Hz}^{-1} \). This corresponds to radio luminosities just below those probed in this analysis, as can be seen in Figure 15, and would imply a two-population model with the more RQ population dominating at \( R \sim 1 \) and below, although without a “dip” at \( R \sim 10 \).

Mahony et al. (2012) claim no bimodality in the \( R_{20} \) radio-loudness parameter with the X-ray-selected sample with \( z < 1 \) down to \( \sim 0.2 \text{ mJy} \). The Kimball et al. (2011) result disfavors a large population with values of \( R \) below \( R < 0.1 \) that would cause a bimodality in the distribution of \( R \), based on a volume-limited sample probed to low radio fluxes. Both of these works claim evidence for a low radio flux population of quasars where the radio emission is significantly enhanced or even dominated by star formation activity in the host galaxies. This is largely a different issue than the “traditional” bimodality question in quasars, i.e., as to whether there is a two-component shape to the radio-loudness distribution centered around \( R = 10 \).

The measured \( dN/dS \) distribution of extragalactic radio sources at low fluxes, together with the limiting level of the extragalactic radio background (Singal et al. 2010; Fixsen et al. 2011) and the quasar optical LF as constructed here, could in principle provide further constraints on the distribution of the radio-loudness parameter at very low values of \( R \). This will be addressed in a forthcoming analysis.

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APPENDIX

NOTES ON THE EFFECTS OF CORRELATED RADIO AND OPTICAL LUMINOSITIES FOR THIS ANALYSIS

As discussed in Section 3.2, in the case of intrinsically correlated luminosities, the LFs in different wavebands are not a priori separable, and one must do a coordinate rotation of the form in Equation (6) to a correlation-reduced luminosity to achieve a separable luminosity function. In that case, one must determine the two now independent luminosity–redshift correlation functions \( g_{\text{opt}}(z) \) and \( g_{\text{cr}}(z) \) that describe the luminosity evolutions. The full procedure is detailed in Section 4, and this result can be compared to the evolutions determined from assuming no intrinsic (i.e., only redshift induced) correlation between the radio and optical luminosities, to determine whether the effect is significant.

The local LFs of uncorrelated luminosities \( L'_{\text{opt}} \) and \( L'_{\text{cr}} \) can then be used to recover the local radio LF by a straightforward integration over \( L'_{\text{cr}} \) and the true local optical LF as

\[
\psi_{\text{rad}}(L'_{\text{rad}}) = \int_0^\infty \psi_{\text{opt}}(L'_{\text{opt}}) \psi_{\text{cr}} \left( \frac{L'_{\text{rad}}}{L'_{\text{opt}}} \right)^{\alpha'} \frac{dL'_{\text{opt}}}{L'_{\text{opt}}} \frac{1}{dL'_{\text{cr}}/dL'_{\text{opt}}} = \psi_{\text{cr}}(L'_{\text{rad}})^{\alpha'} \int_0^\infty \psi_{\text{opt}}(L'_{\text{opt}}) \frac{dL'_{\text{opt}}}{L'_{\text{opt}}} \frac{1}{dL'_{\text{cr}}/dL'_{\text{opt}}}. \quad (A1)
\]
As stated above, this procedure can be used for the determination of the radio LF at any redshift, from which one can deduce that the radio luminosities also undergo luminosity evolution with
\[
g_{\text{rad}}(z) = g_{\text{crr}}(z) \times [g_{\text{opt}}(z)]^{\alpha}
\]  
(cf. Equation (6)).

Similarly, we can determine the local distribution of the radio-to-optical luminosity ratio, \( R' = L'_{\text{rad}} / L'_{\text{opt}} = L'_{\text{crr}} \times L'_{\text{opt}}^{-\alpha} \times L_{\text{fid}}^{-\alpha} \), as

\[
G_R = \int_0^\infty \psi_{\text{opt}}(L'_{\text{opt}}) \psi_{\text{crr}} \left( \frac{R' L_{\text{fid}}}{(L'_{\text{opt}}/L_{\text{fid}})^{\alpha-1}} \right) \frac{dL'_{\text{opt}}}{L'_{\text{opt}}^{\alpha-1} L_{\text{fid}}} 
\]  

(A3)

and its evolution

\[
g_R(z) = g_{\text{crr}}(z) \times [g_{\text{opt}}(z)]^{\alpha-1} = \frac{g_{\text{rad}}}{g_{\text{opt}}}.
\]  

(A4)

These can then be compared with \( \psi_{\text{rad}}(L_{\text{rad}}) \) and \( G_R \) as determined from assuming no intrinsic correlation between the radio and optical luminosities, which is done in Figures 15 and 16. We see that the main conclusions of this work are not dependent on whether the observed correlation between \( L_{\text{rad}} \) and \( L_{\text{opt}} \) is intrinsic or induced.

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