Households’ willingness to pay and preferences for improved cook stoves in Ethiopia

Mekonnen Bersisa1 · Almas Heshmati2 · Alemu Mekonnen3

Received: 9 October 2020 / Accepted: 3 June 2021 / Published online: 12 June 2021
© The Author(s) 2021

Abstract
This paper examines households’ preferences, willingness to pay, and determinants of adopting improved cook stoves in rural Ethiopia. The study uses primary household data selected randomly from three districts in Ethiopia’s Oromia region. The data was collected using a mix of contingent and choice experiment methods of valuation. The former used a double-bounded value elicitation method, while the latter used a fractional factorial design to efficiently generate an attribute and level combination for the improved cook stoves. The study also used various discrete choice models for data analysis and also used models which account for scale and preference heterogeneity. The findings show that the sample households were aware of the effects of using traditional cook stoves and the benefits of using improved cook stoves. However, they were constrained by the availability of the new technology and discouraged by the low-quality of the products that they had used so far. The estimated mean willingness to pay ranged from about 150 Birr to 350 Birr which is lower than the market price of the improved cook stoves. Emission reduction, reducing fire risks, and the durability of the cook stove positively affected its adoption, while price discouraged its use. Higher levels of education, higher incomes, non-farm employment, and having more livestock increased the probability of adopting the new gas stoves. The study recommends that policymakers and product designers should use the mean willingness to pay and marginal rate of substitution for the different attributes as a benchmark for product design and pricing that fit households’ preferences and ability to pay. The lower mean willingness to pay means that a public subsidizing policy is needed for effectively disseminating improved cook stoves in rural Ethiopia.

Keywords Contingent valuation · Choice experiment · Cook stove · Energy-efficient technology · Ethiopia

JEL Classification C25 · D12 · Q51

Introduction

Energy is an important developmental tool at the forefront of the global economic and political agenda. Global environmental problems are largely related to energy use at different levels. Consequently, we observe growing efforts directed towards formulating and intervening in energy policies. Maintaining energy security, expanding access to renewable energy, disseminating energy-efficient technologies, and improving energy use efficiency are some of these policy interventions. However, the effectiveness of any policy intervention depends on societal readiness and support for the intervention (Ruiz-Mercado et al. 2011; Vigolo et al. 2018).

Expanding access to modern energy is tantamount to liberating 2.7 billion people globally who rely on inefficient and traditional energy sources such as firewood, charcoal, dung, and crop residuals for their energy needs. Access to affordable
and reliable energy services in developing countries is fundamental for reducing poverty, improving health, increasing productivity, reducing environmental problems, and promoting economic growth. Modern energy and efficient energy use technologies have multiplier effects for development. They play substantial roles in the provision of clean water, sanitation, and healthcare and provision of reliable and efficient lighting, heating, cooking, and transport and telecommunication services. Despite this, in recent years, nearly 1.2 billion people (about 16% of the global population) in poor countries lacked access to electricity. Most of this electricity-deprived population lived in sub-Saharan Africa (SSA) and South Asia. The country has a target of attaining lower middle-income status by 2025, and its growth path is aligned with the Climate-Resilient Green Economy (CRGE) strategy. Under this plan, the country has embarked on expanding its modern energy sources, and the energy sector is considered an important pillar for realizing green growth and accelerating development. Through its CRGE strategy, Ethiopia has shown the importance of addressing households’ energy demands including expanding renewable energy sources and promoting clean energy technologies (for example, dissemination of 31 million efficient cook stoves by 2030). Promoting efficient stoves is a part of fast-track initiatives for reducing greenhouse gas emissions and emissions from deforestation and forest degradation (FDRE 2011a, 2011b). Nevertheless, the effectiveness of any policy intervention largely depends on consumers’ preferences and willingness to pay for the improvements, their adaptation to a new environment, and continuous innovations in cook stoves and their dissemination for which policy interventions are needed (Ruiz-Mercado et al. 2011; Vigolo et al. 2018).

Vast literature evaluates households’ preferences and willingness to pay for energy-related interventions, but there are limited studies on these issues in developing countries. The few studies that are available focus on willingness to pay for electricity connections, improved cook stove adoption, and different types of fuel in Kenya, India, and Ethiopia (Abdullah and Jeanty 2011; Bhojvaid et al. 2014; Kroon et al. 2014; Takama et al. 2011). Studies in Ethiopia are limited to urban areas, and they are methodologically unable to account for preferences and scale heterogeneity in estimating willingness to pay for improved cook stoves. The problem is less studied in rural areas where a majority of the households in developing countries reside with very different livelihood setups. Besides, people make purchase decisions for a commodity on the basis of some of its features. In this case, the decision-makers claim “attribute non-attendance.” Attribute non-attendance is when respondents ignore a given attribute and its associated level while evaluating an alternative package of attributes. These are behavioral responses and ignoring them in an analysis will bias the estimated mean willingness to pay (Campbell et al. 2011; Hensher and Greene 2010).

The main contributions of this study to existing literature include extending the study to rural areas, incorporating more attributes of improved cook stoves, using a combination of a choice experiment and a contingent valuation method, examining attribute non-attendance, testing different tools for reducing hypothetical bias, and employing discrete choice models which account for preference heterogeneity and scale heterogeneity. In doing so, the study addresses the following research questions: (i) to what extent are rural households aware of the negative effects of using traditional cook stoves and their preference for improved cook stoves?, (ii) what are the cook stove-specific attributes and the socioeconomic determinants of adopting these improved stoves?, (iii) can a

---

1. http://www.worldenergyoutlook.org/resources/energydevelopment/energyaccessdatabase (accessed on 30 April 2020).
2. https://energypedia.info/wiki/Ethiopia_Energy_Situation (accessed on 30 April 2020).
households exhibit preference heterogeneity in choosing the improved cook stove?, and (v) does a household show attribute non-attendance in choosing an improved cook stove?

To address these questions, the study focuses on examining the preferences, willingness to pay, and determinants of the use of improved cook stoves among rural households in Ethiopia thus providing relevant information for rural energy planning and policy.

The rest of this paper is organized as follows. The “Literature review” section presents a review of existing studies on the theory of non-market valuation techniques, cost-benefit analyses, the non-market valuation method, and a theoretical framework of the choice models. It shows the gap in literature on the valuation of improved cook stoves and willingness to pay estimations. The “Data and estimation methodology” section gives details of the data and estimation methodology used. The “Results and discussion” section discusses the results of the study, while the final section gives a conclusion and policy recommendations.

**Literature review**

In reviewing the literature, the focus is on the theory of non-market valuation and cost-benefit analysis, theoretical framework and valuation techniques, and empirical literature on issues related to energy use, adoption of energy-efficient technologies, and households’ preferences and willingness to pay for energy-efficient technologies.

Non-market valuation methods have become an important tool for valuing non-marketed resources (Haab and McConnell 2002; Hanemann et al. 1991). These methods have also been extended to the valuation of marketed goods and services where the conventional market fails to reflect their true values. When there is market failure, market price provides a wrong signal about the economic value of a good or service. The market fails due to the existence of the perverse effect of production and consumption, information asymmetry, lack of well-defined property rights, and existence of public goods (Haab and McConnell 2002; Perman et al. 2003). On the other hand, no market exists for some goods and services making it difficult, if not impossible, to estimate the economic value of such goods. Often, many analytical approaches of project evaluation require some considerations for estimating values in terms of costs and benefits. A peculiar aspect of the cost-benefit analysis is that it requires that the advantages and disadvantages of a project be reduced to numbers which complicates the valuation. Costs of any project are easier to estimate.

A more daunting task is estimating economic benefits. Economists’ non-market valuation techniques are a strong tool for addressing this problem (Ackerman and Heinzerling 2002; Haab and McConnell 2002; Perman et al. 2003).

Non-market valuation methods are generally categorized under two methods: stated preference and revealed preference methods. In stated preference methods, we ask people what they would like to pay or accept for a particular change to happen or not to happen. It is a direct method where people are asked to state their willingness to pay or their choice for a given proposed change. Stated preference methods can be used for preference/choice evaluation, demand analysis, and forecasting. The most frequently used methods here are contingent valuation, choice experiment, contingent ranking, and contingent rating (Haab and McConnell 2002).

Revealed preference methods are indirect non-market valuation methods. These are used for deriving economic value for non-market goods/services indirectly from individual decisions. They encompass travel costs, hedonic pricing, hedonic wages, and avverting behavior methods. The contingent valuation method asks a sample of individuals about their willingness to pay for a hypothetically designed product (Carson 2012; Hanemann 1994). A choice experiment requires a sample of the respondents to make a series of choices from experimentally designed choice sets from which trade-offs between attributes and the marginal rate of substitution are estimated. Applications of choice experiment include economics of the environment, health, transport, marketing, and energy sectors (Adamowicz et al. 1998; Aizak and Nishimura 2008; Haab and McConnell 2002; Hensher et al. 2005).

Non-market valuation techniques are methods of attaching total economic value to goods or services. A non-market valuation exercise is grounded in the theory of welfare economics (Haab and McConnell 2002). Respondents evaluate the effects of any intervention, whether it is an improvement or deterioration on their welfare. An analytical tool that can be used for this is computing the welfare changes of such an intervention. One of the commonly used stated preference valuation techniques is choice experiment. A choice experiment involves following a choice modeling approach which is grounded in the traditional microeconomic theory of choice and decision theory (Hanley et al. 2001; McFadden 1986).

The choice experiment approach combines the characteristic theory of value and the random utility theory (Lancaster 1966; McFadden 1974). The characteristic theory of value states that a consumer’s utility for a good is decomposable into the utility of the characteristics or attributes of the good (Hanley et al. 2001). Moreover, the utility derived from each alternative is assumed to be determined by preferences over the level of attributes provided by that alternative. The assumption that individuals derive utility from the characteristics of a good rather than from the good itself implies that a change in one of the characteristics (such as the price) may result in a discrete change from one good to another which will affect the probability of choosing a specific commodity on the margin (Hanley et al. 1998; Lancaster 1966).
In a choice experiment, a respondent is asked to choose the most preferred among a set of alternatives. Hence, the random utility theory is appropriate for modeling the choices as a function of attributes and attribute levels. In the random utility theory, an individual is assumed to make choices based on the attributes of the alternatives with some degree of randomness. The random utility theory says that the utility derived by individuals from their choice is not directly observable, but an indirect determination of preferences is possible and it decomposes the utility (U) function into a deterministic (V) and a stochastic part (ε). The stochastic part is assumed to follow a predetermined distribution (Brown and Walker 1989; Hanley et al. 2001; McFadden 1974).

There is vast empirical literature on issues related to energy use, adoption of energy-efficient technologies, and households’ preferences and willingness to pay for energy-efficient technologies. Since energy plays a crucial role in economic development and improving social welfare, it has been an area of active inquiry. Most of the studies available so far focus on the effectiveness of interventions, factors affecting energy use technology adoption, and the nexus between energy, poverty, and the environment (Akpalu et al. 2011; Alem et al. 2014; Amigun et al. 2011; Bersisa 2017). Existing literature also shows that individuals’ preferences and tests matter for the effectiveness of policy in promoting energy-efficient technologies. A considerable amount of available studies use stated preference methods for estimating the value of the improvements.

Using the contingent valuation methodology, Abdullah and Jeanty (2011) analyzed rural households’ willingness to pay (WTP) for electricity connections in Kenya. Their study used a double-bounded elicitation format to generate data from the sampled households. They used two econometric methods (parametric and non-parametric) for estimating WTP for an electricity connection and found that the mean WTP of the non-parametric estimation was lower than the mean WTP in the parametric approach. Furthermore, their study showed that respondents’ income, education level, interest in business, age, family size, and house ownership significantly affected WTP for an electricity connection. The study concluded that rural electrification programs should follow a bottom-up approach accounting for consumer preferences and the need for policy formulation.

Similarly, Takama et al. (2011) analyzed households’ willingness to pay and preferences for improved cook stoves in Ethiopia, Tanzania, and Mozambique. Their study emphasized the role of product-specific attributes (indoor smoke, safety, usage cost, and price of the stove) along with socio-economic factors which are commonly used in literature as determinants in studies on improved cook stoves. Takama et al.’s study used the stated preference survey to examine household-level preferences for cooking fuels and stoves. Geographically, the study focused on the capital cities of the three countries to draw sample households: 200 from Addis Ababa, 564 from Dar es Salaam, and 402 from Maputo. Data generated using the stated preference survey was analyzed using a discrete choice analysis for examining the trade-off between the attributes of improved cook stoves. The study concluded that product-specific attributes were as important as households’ socio-economic characteristics and should be given due emphasis in program design, developing products, and policymaking.

A comprehensive study by Bhojvaid et al. (2014) in rural India showed that households mean willingness to pay for improved cook stoves varied significantly according to stove-related attributes. Kooser (2014) examined Ethiopian households’ preferences and willingness to pay for alternative cook stove technologies in Kenya. Kooser (2014) examined Ethiopian households’ preferences and willingness to pay for improved cook stoves. This study showed that adoption of improved cook stoves was affected by various cook stove-specific attributes and socioeconomic characteristics.

Jagger’s and Jumbe’s (2016) study in rural Malawi showed households’ preferences and willingness to adopt a locally produced improved cook stove. Their study used the discrete choice experiment on 383 households randomly selected in rural Malawi for getting their preferences for the locally produced improved cook stove or a package of sugar and salt with equivalent value. Their study showed that the availability of large crop residuals, long time devoted to the collection of fuel wood, awareness about the environmental impact of fuel wood, and peer-effect at the village level increased the odds of choosing the improved cook stove, while availability of a large labor force for fuel wood collection and experience with non-traditional cooking facilities decreased the odds of choosing the improved cook stove.

Using data from rural Guatemala, Bielecki and Wingenbach (2014) showed the importance of cultural and social perceptions in adopting improved cook stoves. Beyond the “triple benefits” of the improved cook stove — health benefits, preserving the local ecosystem, and greenhouse gas reduction — the study argued that the adoption of improved cook stoves was also influenced by other benefits of the stove such as lighting, heating, and becoming a social gathering.

Alem et al. (2014) used panel data for examining the adoption and dis-adoption of improved electric cook stoves in urban Ethiopia. Their study used three rounds of the Ethiopian Urban Socioeconomic Survey for examining the determinants of electric cook stove adoption and dis-adoption which shed light on the state of energy transition in the country. They found that the price of electricity, price of fire wood, and access to credit significantly affected the adoption of electric stoves. However, the study did not include product-specific attributes as determinants of cook stove choice.
Benka-Coker et al. (2018) used different methods for examining the effectiveness, uptake, and scale-up potential of the ethanol Clean Cook stove in a refugee camp and urban settings in Ethiopia. Using the mixed secondary and qualitative primary data, their study applied the Reach, Effectiveness, Adoption, Implementation, and Maintenance (RE-AIM) framework for evaluating the effectiveness and sustainability of the ethanol cook stove intervention and the effectiveness of subsidized distribution of the stove for low-income groups. Their results showed that the improved cook stove had numerous benefits. The study concluded that there were complexities in promoting a new fuel for household cooking, existence of many obstacles, and stagnation in implementation. The study also showed that there was potential for scale-up and commercialization of the ethanol Clean Cook stove in Addis Ababa. It also showed the necessity of stabilizing ethanol supply, providing city-wide distribution infrastructure, and an affordably priced stove and fuel.

Similarly, Mamuye et al. (2018) did a study examining the emissions and fuel use performance and determinants of adopting improved cook stoves in Dodola, southeastern Ethiopia. Their study showed that the improved cook stove reduced CO, CO$_2$, fine particulate matter, and time used for cooking compared to the traditional stove. The study also showed that household head’s sex, age, education level, and income were the determinants for the adoption of the improved cook stove. However, the study did not consider stove-specific attributes that could influence the adoption of the cook stove.

Vigolo et al. (2018) did a systematic literature review for identifying the drivers and barriers to improved cook stoves from a consumer behavior perspective. Their review identified seven categories of determinants of adoption of an improved cook stove: awareness of the risks associated with traditional cook stoves and the benefits of the improved cook stove; attitude towards the technology; economic factors; fuel availability; location; socio-demographics; and social and cultural influences. The study concluded that for the effectiveness of the improved cook stove technology adoption, policymakers and managers should pay due attention to the local context and its social and cultural dynamics.

A handful of studies have applied the stated preference methodology for estimating households’ willingness to pay for improved cook stoves. However, these studies have arrived at conflicting results. For instance, a study by Jeuland et al. (2015) using data from rural households in North India shows that households had strong baseline preferences for the traditional cook stove which will be an inhibiting factor for a wider adoption of the improved cook stove. The study recommends the need for a reinvigorated supply chain with complementary infrastructural investments, appropriate incentives for consumers, and continued applied research and knowledge generation for scaling up the distribution of the improved cook stove.

In sum, existing literature on the adoption of improved cook stoves shows the importance of stove-specific attributes. It also shows that a mix of these attributes greatly affects the adoption of improved cook stoves in addition to the socioeconomic characteristics of the adopters. However, the problem is less studied in rural areas. Moreover, to the best of our knowledge, non-attendance of one or more attributes in making a choice for an improved cook stove has not been studied in Ethiopia. Attribute non-attendance, where respondents ignore a given attribute, and its associated level while evaluating alternative packages of attributes are behavioral responses and hence ignoring them in an analysis will bias the estimated mean willingness to pay (Campbell et al. 2011; Hensher and Greene 2010; Scarpa et al. 2009). This paper contributes to existing literature by extending this study to rural areas, incorporating more attributes of the improved cook stove which fit in rural housing realities and testing the existence of attribute non-attendance. It methodologically uses a mix of different models to account for preferences and scale heterogeneity in choice decisions.

Data and estimation methodology

This section first describes the data and then the estimation methodology. Primary data was generated using a survey of households by employing random sampling technique and the stated preference methodology. The study also used the choice experiment technique for a hypothetically designed improved cook stove scenario. The estimation methodology includes use of different discrete choice models for estimating WTP and its distribution. The marginal effects of attributes and mean WTP for improved cook stoves using multinomial logit model are computed. Determinants of WTP from contingent valuation method are identified and their effects on WTP estimated.

Description of the survey techniques and experimental design

This study was conducted in three selected zones of the Oromia national regional state which is one of the nine regional states in Ethiopia. The selected zones are in the center of the country and in the central part of Oromia region. The selection of these three zones was based on maintaining the heterogeneity of the respondents. These selection criteria were used to account for respondents with different levels of access to forest resources which are important for biomass fuel use. The selected zones are North Showa, South West Showa, and the Finfinne Surrounding Oromia Special Zone. One district was randomly selected from each selected zone. The selected districts are Wachale, Bacho, and Sebete'hasaws. Eight kebeles

---

3 Kebele is the smallest administrative unit, equivalent to a neighborhood with a minimum of 500 households.
were selected from these districts using cluster sampling to include people with different socioeconomic, demographics, and agro-climatic conditions in the sample.

Primary data was generated using a survey of 307 households by employing a multi-stage random sampling technique. Random sampling and sampling proportion to size were used for selecting the households from the selected kebeles. To collect information from selected households, face-to-face interviews were conducted with selected respondents. However, before doing the final survey, a pilot study was done for evaluating the appropriateness and clarity of the instruments and for determining the initial bid (price) used in the final survey. Feedback from the pilot study was incorporated in the final instrument. Five well-trained and experienced enumerators, closely monitored by two supervisors, were involved in the data collection process. The anonymized data was encoded using the SPSS software, and it was then transferred to different versions which are compatible with Stata and Nlogit used for the data analysis in this paper.

The data was collected using the stated preference methodology. It used two stated preference techniques: contingent valuation and choice experiment. For contingent valuation, a scenario in a carefully structured hypothetical market for an energy-efficient cook stove was designed, and households’ willingness to pay for such a product was generated using the double-bounded value elicitation format. This elicitation format was preferred as it fits marketing realities in developing countries where bargaining is common. A respondent was provided with a randomly selected initial bid (proposed price of an improved cook stove) with three bid levels of 75, 150, and 250 birr which was determined from the pilot survey and expert consultations. Depending on the response to the initial bid, a yes or no, the next bid was increased (decreased) by 50%. Following the final response (yes or no), the respondents were asked to state their maximum willingness to pay for the proposed improved cook stove.

The study also used the choice experiment technique for a hypothetically designed improved cook stove scenario. The choice experiment was used for identifying households’ preference for the different attributes of energy-efficient stoves. The respondents were asked to make a series of improved cook stove choices with different combinations of attributes which were mutually exclusive. Detailed information on both attribute-specific and socioeconomic characteristics was generated. In designing a choice experiment, one should consider various factors which can affect the choice of any specific cook stove. These factors can be categorized as stove-specific characteristics which include durability, ease of use, heat energy delivered, start-up time, the price of the stove, and the convenience of use. The second category is related to the risks associated with the use of the technology such as risk of explosion and increasing air pollution levels. The final category includes reliability of use due to different constraints like sustainable supply or availability of energy. Different types of energy-efficient cook stoves exist in Ethiopia albeit with low penetration rates. Mirt, Tikikil, Lakechi, Gonze, and the traditional three-stone cook stoves were considered for the design of the experiment. Of the cook stoves available in the country, the study primarily focused on the cook stove used for baking Enjera since it accounts for 50 to 60% of a household’s energy use (Bizzarri 2010).

Stated preference exercises involve a series of steps. One vital step in a choice experiment survey is determining the number of alternatives, attributes, attribute levels, and values to be considered. Traditionally, a stated preference survey relies on a binary choice for reducing the burden of choice making. Even though it is simple to handle, the binary choice approach has less policy relevance and pragmatism. Thus, extending the number of alternatives and attribute levels improves response quality; it increases the realism of the responses and has room for masking the aim of the study to avoid strategic bias (Hanley et al. 2001).

Hence, for estimating households’ willingness to pay and preferences for the improved cook stove in rural Ethiopia, different attributes of the improved cook stove were selected based on a literature survey, focus group discussions, pilot study, and expert consultations. Various attributes of the improved cook stove were identified. However, due to the complexity of the experimental design and households’ limited cognitive abilities in making choices, we limited the attributes included in the experiment to reducing indoor smoke/emissions, reducing risks of a fire, saving fuel, and durability and price of the improved stove. The proposed improvements to the cook stove were done in such a way that it became a multipurpose stove and could be used for baking Enjera, cooking wat, and boiling coffee. So far, this model has not been developed and distributed in Ethiopia. But making the stove multipurpose was meant to increase energy use efficiency and increase economies of scope and scale advantages in fuel wood use (see Fig. 1).

Along with selecting the attributes of the improved cook stove, it was also necessary to determine different levels corresponding to each attribute. Three levels were selected for each attribute of the improved cook stove. The starting points for level selection for each attribute were driven by empirical studies on the topic (Beyene et al. 2015). The attributes are as follows:

1. **Indoor smoke or emission reduction:** indoor smoke is an acute health problem that rural households face because they rely on traditional cook stoves and traditional energy sources. High indoor emission levels expose household members to different health risks. Thus, this program is designed to reduce the problems of indoor smoke because
of using the traditional three-stone cook stove. The program integrates different packages with the improved cook stove to reduce emission levels by 40, 65, and 90%.

2. Risk of use: this program was designed in response to the different hazards of using the traditional cook stove. It involves designing a cook stove which reduces risks of a fire, burning, and explosions. It makes households safer and reduces related costs for prevention and curatives. Risk level is defined as low, medium, and high.

3. Fuel saving: shortage of fuel wood and higher costs of fuel and its environmental effects through deforestation are some of the problems related to fire wood collection and use. This program is meant to improve the stove’s efficiency and reducing fuel wood usage by 25, 45, and 65%.

4. Durability of the stove: frequent breakages of the stove expose a household to unnecessary costs. This program aims at producing a stove which can serve a household for different durations — 5, 15, or 20 years — depending on the material used in its production.

5. Price of the stove: it is unquestionable that price affects the demand for a product. Three price levels of 75, 150, and 250 birr were used in the survey to examine the effect of price on the demand for the stove and for the computation of marginal willingness to pay for each attribute (Table 1).

Five attributes with three levels each were used for constructing the choice sets. The algorithm of experimental design of the R software in the discrete choice experiment was used for the experimental design and for testing the efficiency of the design. The SAS software was used for supplementing the experimental design and for testing the efficiency of the choice experiment’s design. With a D-efficiency of 97.8%, the orthogonal fractional factorial design was developed for the first alternative, reducing the original full factorial design of 243 (3⁵) possible combinations to 18. It was not feasible to present the full factorial design with all the possible treatment combinations as choice tasks. As a result, only a fraction of the total number of treatment combinations was selected. Optimum treatment combinations were selected such that all the attributes were statistically independent of each other (orthogonal). Finally, the 18 treatment combinations selected were divided into three versions to minimize inefficient selection in multiple choice tasks. The discrete choice experiment provided a panel of six choice sets for a given version (see Table 2). Each choice set consisted of two experimentally designed alternatives — labeled “option 1” and “option 2” — and a status quo alternative — labeled “no action” — which portrayed all the attributes of the currently in-use three-stone cook stove with no additional payment.

Table 1 Attributes and levels of the proposed improved cook stove

| Attributes/stove type                  | Levels                      | Traditional stove |
|---------------------------------------|-----------------------------|-------------------|
| Indoor smoke/emission reduction       | 40, 65, and 90%             | Status quo        |
| Risk of use                           | Low, moderate, high         | Status quo        |
| Fuel wood saving                      | 25, 45, and 65%             | Status quo        |
| Durability of the stove (in years)    | 5, 15, 20                   | Status quo        |
| Price of the improved stove (in birr*)| 75, 150, 250                | Status quo        |

Source: Developed by the authors

*US$1 = 35.06 Ethiopian birr (ETB) on 7 August 2020.
As we can see in Table 2, understanding improvements in each attribute in the choice sets could be difficult for some respondents. As the survey was conducted in rural areas where a majority of the respondents have low educational levels, it was necessary to supplement the choice sets by a choice card which diagrammatically represented the choice sets. To assist respondents and to make the choice task as simple as possible, choice cards were prepared for each choice set and presented to the respondents. A sample choice card is given in Fig. 2.

Table 2 Sample choice sets used for the survey. Source: Designed by the authors.

| Attributes                  | Option 1   | Option 2   | Option 3   |
|-----------------------------|------------|------------|------------|
| Emission reduction          | 90%        | 40%        | Status quo |
| Fuel saving                 | 45%        | 65%        | Status quo |
| Risk of fire                | Low        | Moderate   | Status quo |
| Durability of the stove     | 5 years    | 15 years   | Status quo |
| Price of the stove (in Birr)| 250        | 150        | Status quo |
| Which one will you choose?  |            |            |            |

Model specification and estimation methods

In this section, we introduce discrete choice models of multinomial logit, mixed multinomial logit, and scale heterogeneity multinomial logit that are used for analyzing choice experiment data. The analysis of the improved cook stove choices is further extended to the generalized mixed logit model. Models for estimating WTP distribution using analysis of the double-bounded response data from the contingent valuation method survey are introduced which uses interval data. The optimal class size is selected, and marginal WTP for improved cook stove’s attributes using multinomial logit model is computed. Determinants of WTP are identified and their effects on WTP estimated. Finally, mean WTP using different models is estimated. The multinomial model and its generalization are described below.

Models used for analyzing choice experiment data

Models of discrete choice data are grounded in the theoretical underpinnings of the characteristic theory of value (Lancaster 1966) and the random utility theory (McFadden 1974). An individual chooses an improved cook stove since he/she values the attributes of the product. Here, we assume utility as a latent construct that underlies observed choices reflecting the demand for a good. Respondent n is assumed to consider the full set of offered alternatives i and choosing the alternative with the highest utility. As implied by the characteristic theory of value, utility of option i for individual n (Uin) is assumed to depend on the attributes (Zi) of the good to be valued and the socioeconomic characteristics of individual users (Sn). This utility is decomposed into deterministic and stochastic components as:

\[ U_{in} = V_{in} + \varepsilon_{in} \]  

(1)

where Uin is the latent, unobserved utility of consumer n for choice alternative i, V_{in} is the deterministic part of the utility that individual n has for choice alternative I, and \( \varepsilon_{in} \) is the random portion of the utility that consumer n has for choice alternative i. Employing the rationality assumption, that is, individuals are utility maximizers, the probability that individual i will choose option j over j is given by:

\[ \text{Prob}(i|R) = \text{Prob}(U_{in} > U_{jn}) = \text{Prob}(V_{in} + \varepsilon_{in} > V_{jn} + \varepsilon_{jn}) \]  

for \( j \in R \) and \( i \neq j \)  

(2)

where R is the complete choice set available to the individual. This probability is estimable under the assumption that the error terms are independently and identically distributed with extreme-value distribution. This assumption gives rise to the specification of the multinomial logit model (MNL) that determines the probabilities of choosing alternative i over j (Hanley et al. 2001):

\[ \text{Prob}(U_{in} > U_{jn}) = \frac{\exp(\mu V_i)}{\sum_{j \neq i} \exp(\mu V_j)} \]  

(3)

where \( V_i = V(Z_i, S) \) is the indirect utility function, \( Z_i \) is a vector of the energy-efficient cook stove’s attributes, S is a vector of users’ socioeconomic characteristics, and \( \mu \) is a scale parameter inversely related to the standard deviation of the error term. The implication of this is that the estimated \( \beta \)s cannot be directly interpreted for their contribution to utility since they are confounded with the scale parameter. The MNL model must satisfy independence from irrelevant alternatives (IIA) properties, which means that the addition or subtraction of any alternatives to the choice available to the respondents will not affect the
relative probability of individual \( n \) choosing any other alternative (Hausman and McFadden 1984). The deterministic utility function from the MNL model can be presented by \( V_{ij} \) which is assumed to have an additive structure and is given by:

\[
V_{ij} = \text{ASC} + \sum \beta_k Z_k + \sum \beta_m S_m \quad (4)
\]

where ASC (alternative-specific constant) captures systematic variations in choice observations which are associated with alternatives that are not explained either by an attribute’s variations or by respondents observed socioeconomic characteristics (Ben-Akiva and Lerman 1985). It accounts for the effects of any attribute not included in the choice set for utility (Agimas and Mekonnen 2011). \( \beta_k \) is a vector of coefficients corresponding to \( Z \)’s attributes from the choice sets, and \( \beta_m \) is a vector of coefficients corresponding to \( S \)’s socioeconomic characteristics of the respondents. To improve the efficiency of the estimated coefficients and WTP, the study used the bootstrap method (Cameron and Trivedi 2010; Takama et al. 2011). Once the multinomial logit model’s estimations were obtained, the marginal value of each attribute and the corresponding marginal rates of substitution were estimated. Besides, a measure of welfare change (compensating surplus) that conforms to the demand theory can be derived from the proposed changes in the improved cook stove’s attributes. The marginal rate of substitution for each attribute is estimated following Hanley et al. (2001):

\[
\text{WTP} = b_y \ln \left( \frac{\sum_i \exp \left( V_{i1} \right)}{\sum_i \exp \left( V_{i0} \right)} \right) \quad (5)
\]

where \( V_{i0} \) represents the utility of the initial state and \( V_{i1} \) represents the utility of the alternative state. The parameter \( b_y \) is the coefficient of price attribute which measures the marginal utility of income. For the linear utility index representation, WTP is simply the ratio of the coefficient of an attribute to the coefficient of payment (price). This ratio is called part-worth/implicit price or marginal willingness to pay for the attribute. It is estimable in Nlogit using the Krinsky and Robb method (Hensher et al. 2005):

\[
\text{WTP} = \frac{\beta_{\text{attribute}}}{\beta_{\text{payment}}} \quad (6)
\]

The assumption of IIA in the MNL model is hard to meet in real choice models. If assumption of IIA is violated (which can be tested using the Hausman and McFadden (1984) procedure), we have to resort to models which relax this assumption. One such model is the random parameter logit (RPL) model also known as the mixed logit model (Train 1998). As opposed to the MNL model, the mixed logit model (MLM) allows for high flexibility by specifying taste coefficients to be randomly distributed across individuals, and it accounts for households’ unobserved
heterogeneity in decision-making (Campbell et al. 2011; Hensher and Greene 2003; Kroon et al. 2014). It also allows for interdependence of the choice situation by allowing the AR(1) process. MLM generalizes a standard MNL by allowing its parameters associated with the observed variable to vary with a known population distribution across individuals. It also assumes continuous joint distribution and is specified as:

\[ P_{ijq} = \frac{\exp(\alpha' + \beta X_{ij} + \varphi F_{ij})}{\sum_j^{J} \exp(\alpha' + \beta X_{ij} + \varphi F_{ij})} \]  

(7)

where \( \alpha' \) is a vector of fixed or random ASCs associated with \( i=1, \ldots, J \) alternatives and \( q=1, \ldots, Q \) individuals and one of these ASCs should be identified as 0. \( \beta' \) is a parameter vector that is randomly distributed across individuals. \( \varphi' \) is a vector of non-random parameters. \( X_{ij} \) is a vector of individual-specific characteristics and alternative-specific attributes at choice situation \( t \) and is estimated with random parameters. \( F_{ij} \) is a vector of individual-specific characteristics and alternative-specific attributes at choice situation \( t \) and is estimated with fixed parameters. As an alternative to the mixed logit model which assumes continuous joint distribution of the sources of individual preference heterogeneity, we also use the latent class model which assumes discrete distribution of preference heterogeneity.

Following Boxall and Adamowicz (2002) and Greene and Hensher (2003), we use a latent class model for analyzing the behavior of different classes in choosing energy-efficient cook stoves. Here, a consumer shows discrete preference heterogeneity for goods and services. This consumer preference heterogeneity must be accounted for while estimating a consumer’s WTP. The latent class model’s specification stands on the theory that individual behavior depends on observable attributes and on latent heterogeneity that varies with factors that are unobserved by an analyst (Greene and Hensher 2003; Shen 2009). This heterogeneity can be analyzed by the model of discrete parameter variation. The central assumption in this model is that individuals are implicitly sorted into a set of \( Q \) classes and the analyst has no prior information about the class to which an individual belongs. Following Greene and Hensher (2003), the latent class model can be extended from the central behavioral model of the multinomial logit model for discrete choices among \( J \) alternatives by individual \( i \) observed in \( T_i \) choice situations:

\[ \Pr(\text{individual } i \text{ chooses } j \text{ in choice situation } t|\text{class } q) = \frac{\exp(X_{itj} \beta_q)}{\sum_{j=1}^{J} \exp(X_{itj} \beta_q)} = \Phi(i, j, t|q) \]  

(8)

Consequently, the probability of a specific choice made by an individual \( i \) given the class to which s/he belongs can be specified as:

\[ P_{i|q}(j) = \Pr(y_{it} = j|\text{class } q) \]  

(9)

In this formulation, for the given class assignment, the contribution of individual \( i \) to the likelihood function will be the joint probability of the sequence \( y_{it} = \{y_{i1}, y_{i2}, \ldots, y_{iT}\} \) given as:

\[ P_{i|q} = \prod_{t=1}^{T_i} P_{i|q} \]  

(10)

The class assignment is not known, but if \( H_{iq} \) denotes the prior probability for class \( q \) for individual \( i \), it can conveniently be presented as a multinomial logit as:

\[ H_{iq} = \frac{\exp(z_i q)}{\sum_{q=1}^{Q} \exp(z_i q)} \]  

(11)

where \( z_iq \)s are a set of observable characteristics which enter the model of class membership. To enable identification of the model, the \( Q \)th parameter is normalized to zero. The likelihood of individual \( i \) is the expectation (over classes) of class-specific contributions given as:

\[ P_i = \sum_{q=1}^{Q} H_{iq} P_{i|q} \]  

(12)

Finally, the log likelihood function for the sample can be specified as:

\[ \ln L = \sum_{i=1}^{N} \ln P_i = \sum_{i=1}^{N} \ln \left[ \sum_{q=1}^{Q} H_{iq} \prod_{t=1}^{T_i} P_{i|q} \right] \]  

(13)

The coefficients are estimated for all classes through maximizing the likelihood function. However, the striking fact here is that the determination of the number of classes for the latent class model is not straightforward. This study examines the optimum class size using the information criteria and comes up with two classes for estimating this model.

Currently, researchers who prefer analysis and discrete choice models have come up with more sophisticated and advanced models which account for both preference heterogeneity and scale heterogeneity. The generalized multinomial logit model (GMNL) is one such model. Following Fiebig et al. (2010), the GMNL model for a sample of \( n \) respondents with the choice of \( J \) alternatives in \( T \) choice situations represents the probability of respondent \( i \) choosing alternative \( j \) in choice situation \( t \) as:

\[ \Pr(\text{choice } i = j|\beta) = \frac{\exp(\beta_j X_{ij})}{\sum_{j=1}^{J} \exp(\beta_j X_{ij})}, \quad i = 1, \ldots, N; \quad t = 1, \ldots, T \]  

(14)

where \( X_{ij} \) is a vector of observed attributes of alternative \( j \), and \( \beta_j \) is a vector of individual-specific parameters defined as:
\[ \beta_i = \sigma_i \beta + \{ \gamma + \sigma_i (1-\gamma) \} \eta_i \]  

(15)

Here, the specification of \( \beta_i \) is central to GMNL. It depends on a constant vector \( \beta \), a scalar parameter \( \gamma \), a random vector \( \eta_i \) distributed multivariate normal, MVN(0, \( \Sigma \)), and \( \sigma_i \), the individual-specific scale of the idiosyncratic error. The value of \( \gamma \) ranges from [0, 1], and at extreme values of \( \gamma \), we get GMNL type I or II. GMNL is able to account for “extreme” consumers with nearly lexicographic preferences. It is able to explain consumers who exhibit very “random” behavior. This paper extends the analysis of the improved cook stove choices to the generalized mixed logit model.

**Estimating WTP distribution**

Estimating WTP for goods and services serves several purposes, and hence, it has been used in different areas. There are several methods for estimating WTP. The first and simplest method is to directly ask the respondents their willingness to pay for the good/service under consideration. However, this method has several problems: respondents may face cognitive difficulties and they may behave strategically while responding to different incentives (Hole and Kolstad 2012). This can be estimated either in the preference (utility space) or in the WTP space. The distribution of WTP in the utility space can also be specified. The utility of household \( n \) derives from the use of cook stove \( j \) under situation \( t \) specified as a function of income of household \( w_{njt} \) and other non-income attributes of the stove \( x_{njt} \) written as:

\[ U_{njt} = \alpha_n w_{njt} + \beta_n X_{njt} + \varepsilon_{njt} \]  

(16)

From Eq. (16), \( \alpha_n \) and \( \beta_n \) indicate individual-specific coefficients for income and the other attributes of the improved cook stove chosen, and \( \varepsilon_{njt} \) is a random error term. We assume that \( \varepsilon_{njt} \) is the extreme value distributed with variance given by \( \mu_n \), where \( \mu_n \) is an individual-specific scale parameter. Train and Weeks (2005) show that dividing Eq. (16) by \( \mu_n \) does not affect the behavior and results in a new error term which is i.i.d. extreme value distributed with variance equal to \( \frac{1}{\mu_n^2} \) and we get a new equation:

\[ U_{njt} = \alpha_n w_{njt} + \frac{\beta_n}{\mu_n} X_{njt} + \frac{\varepsilon_{njt}}{\mu_n}, \]  

(17) \[ U_{njt} = \lambda_n w_{njt} + C_n X_{njt} + \varepsilon_{njt} \]  

(18)

Equation (18) is a specification of WTP in the preference space (Train and Weeks 2005). Moreover, decision-makers are highly influenced by one characteristic of a product during purchase decisions. In the choice experiment, this is known as “attribute non-attendance.” It is an attribute that is not considered by decision-makers or is ignored while making a choice; this has occupied a prominent place in discrete choice models (Hensher et al. 2012). Our study examined the presence of attribute non-attendance in the improved cook stove choice decisions. It used the attribute non-attendance where respondents were asked to state to what extent they had attended to each attribute after the choice exercise was complete.

An analysis of the double-bounded response data from the contingent valuation method (CVM) survey uses interval data. Information which is directly elicited is a dichotomous response taking a value of zero if the individual says no and one if the individual answers yes to a question. Individual’s response depends on the price (bid) for the product with improvements proposed to be provided. Let us assume that an individual is asked if he is willing to pay \( t_i \) for a given change in the good proposed to be provided. If the individual’s response is no, then we can say his willingness to pay is between zero and \( t_i \), that is, \( 0 \leq \text{WTP} \leq t_i \) and if he answers yes, his willingness to pay is \( t_i \leq \text{WTP} < \infty \). For getting an accurate estimate of WTP, we need a relatively larger sample size which is hardly attainable in such a survey.

Alternatively, the double-bounded methodology proposed by Hanemann et al. (1991) can be used for efficiently estimating WTP. In this case, a respondent is given a follow-up question after he responds yes or no to the first bid question. If the individual answers yes to the first question, he is provided with a higher bid. On the other hand, if the individual answers no to the first bid question, then he is offered a lower bid. For examining the individual’s WTP distribution following Lopez-Feldman’s (2012) explanation, we can denote the first bid price by \( t^1 \) and the second bid price by \( t^2 \). In this regard, each respondent will fall in one of the following categories:

- If an individual answers yes to the first bid question and no to the second bid question, then \( t^1 > t^2 \). In this case, we can say that \( t^1 \leq \text{WTP} \leq t^2 \).
- If an individual answers yes to both the questions, then we have \( t^1 \leq \text{WTP} < \infty \).
- If an individual answers no to the first question and yes to the second question, then \( t^1 > t^2 \), and we have \( t^2 \leq \text{WTP} < t^1 \).
- Finally, if an individual answers no to both the first and second questions, we have \( 0 < \text{WTP} < t^1 \).

These four cases can be estimated by the double-bounded or interval data model. To specify the model, let us denote the answer to the first and second response questions by dichotomous variables \( y_i^1 \) and \( y_i^2 \). The probability that an individual will answer yes to the first question and no to the second question is given as \( \Pr(y_i^1 = 1, y_i^2 = 0 | z_i) = \Pr(s, n) \). Specifying this probability distribution works under the assumption that:

\[ \text{WTP}(z_i, u_i) = Z_i \beta + u_i \text{ and } u_i \sim \text{N}(0, \sigma^2) \], then the probability distribution of the four choices is:
Case 1: $y_1 = 1$ and $y_1^2 = 0$

\[ Pr(\text{yes, no}) = \Pr(t_1 \leq \text{WTP} < t_2) = \Pr(t_1 \leq Z_i \beta + u_i < t_2) \]

\[ = \Phi \left( \frac{t_2 - Z_i \beta}{\sigma} \right) - \Phi \left( \frac{t_1 - Z_i \beta}{\sigma} \right) \]

where the last expression follows $\Pr(a \leq X < b) = F(b) - F(a)$. Therefore, using the symmetry of the normal distribution, we have:

\[ Pr(\text{yes, no}) = \Phi \left( \frac{Z_i \beta - t_1}{\sigma} \right) - \Phi \left( \frac{Z_i \beta - t_2}{\sigma} \right) \]  \hspace{1cm} (20)

Case 2: $y_1 = 1$ and $y_1^2 = 1$

\[ Pr(\text{yes, yes}) = Pr(\text{WTP} > t_1, \text{WTP} > t_2) \]

\[ = \Pr(Z_i \beta + u_i > t_1, Z_i \beta + u_i > t_2) \]

\[ = \Phi \left( \frac{Z_i \beta + u_i > t_1}{\sigma} \right) \Phi \left( \frac{Z_i \beta + u_i > t_2}{\sigma} \right) \]  \hspace{1cm} (21)

Using Baye’s rule, which says that $Pr(A,B) = Pr(A|B)*Pr(B)$, we have:

\[ Pr(\text{yes, yes}) = \Pr \left( Z_i \beta + u_i > t_1 | Z_i \beta + u_i > t_2 \right)*Pr \left( Z_i \beta + u_i > t_2 \right) \]  \hspace{1cm} (22)

Here by definition $t^2 > t^1$ and then $Pr(Z_i \beta + u_i > t_1 | Z_i \beta + u_i > t^2) = 1$ which implies

\[ Pr(\text{yes, yes}) = \Pr(u_i > t^2 - Z_i \beta) = 1 - \Phi \left( \frac{t^2 - Z_i \beta}{\sigma} \right) \]

So, by symmetry we have:

\[ Pr(\text{yes, yes}) = \Phi \left( \frac{Z_i \beta - t^2}{\sigma} \right) \]  \hspace{1cm} (23)

Case 3: $y_1 = 0$ and $y_1^2 = 1$

\[ Pr(\text{no, yes}) = \Pr(t_1 \leq \text{WTP} < t_2) \]

\[ = \Pr(t_1 \leq Z_i \beta + u_i < t_2) \]

\[ = \Phi \left( \frac{t_2 - Z_i \beta}{\sigma} \right) - \Phi \left( \frac{t_1 - Z_i \beta}{\sigma} \right) \]  \hspace{1cm} (24)

\[ Pr(\text{no, yes}) = \Phi \left( \frac{Z_i \beta - t_1}{\sigma} \right) - \Phi \left( \frac{Z_i \beta - t_2}{\sigma} \right) \]

\[ \text{Case 4: } y_1 = 0 \text{ and } y_1^2 = 0 \]

\[ Pr(\text{no, no}) = \Pr(\text{WTP} < t_1, \text{WTP} < t_2) \]

\[ = \Pr(Z_i \beta + u_i < t_1, Z_i \beta + u_i < t_2) \]

\[ = \Phi \left( \frac{t_2 - Z_i \beta}{\sigma} \right) = 1 - \Phi \left( \frac{t_1 - Z_i \beta}{\sigma} \right) \]

(25) \[ Pr(\text{no, no}) = \Pr( \text{WTP} < t_1, \text{WTP} < t_2) \]

\[ = \Pr \left( Z_i \beta + u_i < t_1, Z_i \beta + u_i < t_2 \right) \]

\[ = \Phi \left( \frac{t_2 - Z_i \beta}{\sigma} \right) = 1 - \Phi \left( \frac{t_1 - Z_i \beta}{\sigma} \right) \]

For these four models, we have to construct the likelihood function to directly obtain $\beta$ and $\sigma$ through the maximum likelihood estimation method. To be maximized and for finding the parameters of the model, the likelihood function can be constructed as:

\[ \prod_{i=1}^{N} \left[ d_{i,\text{yes, no}}^{\text{WTP}} \Phi \left( \frac{t_1 - Z_i \beta}{\sigma} \right) - d_{i,\text{yes, yes}}^{\text{WTP}} \Phi \left( \frac{t_1 - Z_i \beta}{\sigma} \right) + d_{i,\text{no, no}}^{\text{WTP}} \left[ \Phi \left( \frac{t_1 - Z_i \beta}{\sigma} \right) - \Phi \left( \frac{t_2 - Z_i \beta}{\sigma} \right) \right] \right] \]

\[ \sum_{i=1}^{N} \left[ d_{i,\text{yes, no}}^{\text{WTP}} \Phi \left( \frac{t_1 - Z_i \beta}{\sigma} \right) - d_{i,\text{yes, yes}}^{\text{WTP}} \Phi \left( \frac{t_1 - Z_i \beta}{\sigma} \right) + d_{i,\text{no, no}}^{\text{WTP}} \left[ \Phi \left( \frac{t_1 - Z_i \beta}{\sigma} \right) - \Phi \left( \frac{t_2 - Z_i \beta}{\sigma} \right) \right] \right] \]

where $d_{i,\text{yes, no}}^{\text{WTP}}$, $d_{i,\text{yes, yes}}^{\text{WTP}}$, $d_{i,\text{no, no}}^{\text{WTP}}$, and $d_{i,\text{no, no}}^{\text{WTP}}$ are indicator variables that take values of one or zero depending on the relevant case for each individual whereby each individual will only appear once in the likelihood function. Using this likelihood function, estimating the parameters ($\beta$ and $\sigma$) is straightforward.\(^5\)

Finally, this paper tested one method for reducing a hypothetical bias in the contingent valuation survey. Biases abound in contingent valuation studies. People’s intentions and actions deviate in the real world. These can be strategic bias, hypothetical bias, and starting point bias. Several methods are proposed and implemented in contingent valuation literature to reduce these biases. Hypothetical bias is broadly defined as a difference between stated and revealed WTP. It is a systematic divergence between welfare estimates obtained through the stated preference and revealed preference choice instruments. Usually, WTPs stated by the individuals often exceed their real-money WTPs (List and Gallet 2001; Murphy et al. 2005). This paper used cheap-talk to test its impact on reducing the hypothetical bias in the contingent valuation survey. Cheap-talk is communication between

\(^5\) The Doubleleb Stata command is used for estimating this interval data model after adding the ado developed by Lopez-Feldman (2012) in Stata Version 12.
players that does not directly affect the pay-offs of the game. In such a game, providing and receiving information is free.

**Results and discussion**

**Descriptive statistics**

This section presents some descriptive statistics of the responses to some of the questions in the survey. It also briefly presents respondents’ perceptions, awareness, socioeconomic characteristics, and ownership status. It is aimed to give readers some clues about the nature of the data and its distribution. Table 3 gives summary statistics for some of the continuous variables.

As can be seen in Table 3, on average, the age of the respondents was 42.82 years with a standard deviation of 12.90 years. Household heads in the study area, on average, achieved grade 3, and family size was about six which is relatively higher than the family size in rural Ethiopia. On average, sample households’ monthly income was about 2116 birr, and monthly expenditure was slightly lower than the income at about 1542 birr. Off-farm employment was not common in the study area. Respondents’ land ownership status shows that, on average, a household possessed about 7.27 kert (less than 2 ha) of land. It seems that landholding in the study area did not meet the land demands of a majority of the farmers as they were, on average, net renters-in. Oxen were predominantly used as a source of power for plowing. On average, a household owned two oxen and about 5.25 livestock in terms of tropical livestock units.

**Table 3** Descriptive statistics of continuous variables (n = 307)

| Variables                        | Mean  | Std. dev. |
|----------------------------------|-------|-----------|
| Age of the respondent (in years) | 42.82 | 12.90     |
| Education level of the respondent (in years) | 2.96  | 3.57      |
| Household size (in number)       | 6.35  | 2.44      |
| Adult male equivalence household size | 5.05  | 2.05      |
| On-farm employed members         | 2.10  | 1.49      |
| Off-farm employed members        | 0.14  | 0.37      |
| Land owned in local units (kert) | 7.27  | 6.54      |
| Land cultivated in local units (kert) | 9.39  | 8.23      |
| Land rented in local units (kert) | 3.28  | 6.03      |
| Land rented out in local units (kert) | 0.34  | 1.42      |
| Number of oxen owned             | 2.01  | 1.56      |
| Livestock owned (in TLU)         | 5.25  | 3.91      |
| Household expenditure per month (in birr) | 1541.68 | 980.79    |
| Household income per month (in birr) | 2115.57 | 1661.49  |

Source: Authors’ computations using survey data

Table 4 shows that close to two-thirds of the respondents did not have access to credit. Regarding gender of household heads, about 67.4% of the households were male-headed. The penetration rate of education in rural parts in developing countries is very low. In case one wants to talk about education in these areas, it is more appropriate to talk about literacy rates. About 60% of the household heads in the study area could read and write.

More than half of the respondents had experienced different health problems as a result of their current cook stove (Table 5). Physical burns, respiratory diseases, symptoms of asthma, and irritation in the eyes and nose were some of the problems that they had experienced. The results in Table 5 show that a majority of the respondents (about 95%) had information about the improved cook stove. Their sources of information included health extension workers in the area, agricultural development workers, friends, radio, and local community organizations like Idir and Iqub. Despite this information, the penetration rate of the improved cook stove was very low. Only about 26% of the respondents were using the improved cook stove. The responses to a follow-up

**Table 4** Frequency of categorical variables

| Variables                        | Category | Frequency | Percentage |
|----------------------------------|----------|-----------|------------|
| Access to credit                 | Yes      | 116       | 37.79      |
|                                  | No       | 191       | 62.21      |
| Respondent’s sex                 | Male     | 207       | 67.43      |
|                                  | Female   | 100       | 32.57      |
| Can read and write               | Yes      | 183       | 59.61      |
|                                  | No       | 124       | 40.39      |
| Owns a private tree              | Yes      | 265       | 86.32      |
|                                  | No       | 42        | 13.68      |
| Marital status                   | Married  | 274       | 89.25      |
|                                  | Otherwise| 33        | 10.75      |

Source: Authors’ computations using survey data

**Table 5** Frequency and distribution of use, health effects, and information about the cook stove

| Variables                        | Category | Frequency | Percentage |
|----------------------------------|----------|-----------|------------|
| Have you had any health-related injury when cooking on a traditional stove? | Yes | 195 | 63.52 |
|                                  | No       | 112       | 36.48      |
| Have you heard about the improved cook stove? | Yes | 290 | 94.46 |
|                                  | No       | 17        | 5.54       |
| Are you currently using any improved cook stove? | Yes | 79 | 25.73 |
|                                  | No       | 228       | 74.27      |
| Do you use the same type of stove for cooking all meals? | Yes | 184 | 59.93 |
|                                  | No       | 123       | 40.07      |

Source: Authors’ computations using survey data
question showed that the respondents were conservative in using the product due to a bad experience and also because they had very little information about using the product.

The results in Table 6 give a clear picture of respondents’ perceptions about using the improved cook stove. Use of the traditional cook stove has tremendous negative health effects. Indoor air pollution, physical burning, and damage to different body parts due to fire and smoke are common health effects of using the traditional cook stove. In line with this, households strongly believed that use of the improved cook stove will reduce these problems. About 62.5% of the respondents strongly agreed that the use of the improved cook stove will improve the health of household members. A majority of the respondents also agreed that the improved cook stove saved time. They also agreed that use of the improved cook stove should be mandatory in Ethiopia to alleviate the problems of using the traditional (three-stone) cook stove.

Theoretically, as price increases, demand for a product decreases, other things remaining constant. As presented in Table 7, the proportion of respondents who accepted the offered price of the stove (bid) decreased as its price increased. This is in compliance with the theory of valuation which states that as the bid value increases respondents’ willingness to accept the product decreases.

It is evident from Table 8 that attribute non-attendance was not a serious problem in this choice experiment. A fairly significant number of respondents attended to almost all the attributes while making choice decisions. Fuel wood saving was considered less in the choice exercise. This can be attributed to the fact that rural households are not constrained by the availability of fuel wood.

**Econometric results**

A starting point for discrete choice models, the multinomial logit model (MNL) and the mixed logit model (MLM), is running the basic model with only alternative-specific attributes as explanatory variables (Hensher and Greene 2003). Table 9 presents the results of the basic MNL, MLM, the scale heterogeneity multinomial logit model (SMNL), and the generalized mixed logit model (GMNL). Column 2 of Table 9 gives the results of the MNL model. The overall fit of the model as measured by McFadden’s Pseudo R² is good (0.20), but using the log likelihood ratio test and information criterion, it performs poorly compared to the other models. The coefficients of this model are significant at the less than 1% significance level except for fuel wood saving. This indicates that all the selected attributes except fuel wood saving significantly affected the choice of the improved cook stove.

However, the MNL model is estimated under a stringent assumption of IIA. This study tested this assumption by using the Hausman test by excluding one of the alternatives. The chi-square value of 17.96 with a p-value of 0.006 shows that the IIA assumption was violated in the MNL model. As a result, this study estimated the alternative model which relaxes this assumption. The first model is the random parameter logit model, also known as the mixed logit model. The merits of this model not only are relaxing the assumption of IIA, but it also considers preference heterogeneity. The results of the mixed logit model are presented in column 4 of Table 9. The estimates of standard deviations show that there is preference heterogeneity in choosing the improved cook stove.

The results in Table 9 show that for all the models, the coefficient of an alternative-specific constant is positive and significant. The positive and significant alternative-specific constant (ASC) coefficient indicates that respondents had higher utility for policy alternatives (improved cook stove) compared to the status quo which is similar with the findings of Jagger and Jumbe (2016) but contrasts with the findings of Jeuland et al.’s (2015) study which showed that consumers exhibited a stronger baseline preference for the traditional stove. As expected, the coefficient of price was negative and significant indicating that respondents had higher utility for alternatives with lower price levels. The positive and significant coefficients of emission reduction, reducing risks of use, and the durability of the stove show that the respondents derived higher utility from the stove with lower levels of emissions, lower risks of fire, and longer durability of the stove.

| Table 6 | Households’ perceptions about the improved cook stove |
|---------|--------------------------------------------------------|
| **Variables** | **Strongly agree** (in %) | **Agree** (in %) | **Undecided** (in %) | **Disagree** (in %) | **Strongly disagree** (in %) |
| Use of the improved cook stove improves your family’s health. | 62.54 | 30.62 | 6.84 | 0 | 0 |
| Use of the improved cook stove saves time. | 62.540 | 32.25 | 5.21 | 0 | 0 |
| It is easy to buy the improved cook stove in your area. | 54.40 | 20.85 | 6.19 | 13.36 | 5.21 |
| Fixing the improved cook stove is easy. | 39.41 | 37.46 | 13.68 | 8.47 | 0.98 |
| Use of the improved cook stove should be mandatory in Ethiopia. | 61.24 | 29.97 | 5.86 | 2.93 | 0 |

Source: Authors’ computation using survey data

© Springer
Our findings are similar to that of Takema et al.’s (2011) and Kooser’s (2014) findings. Stove-specific attributes are important in determining the adoption of the improved cook stove. In our study, the fuel-saving feature of the improved cook stove became insignificant in all the models that we estimated which is different from Kooser’s (2014) findings. This might be due to cultural and multiple benefits of firewood use in rural parts as besides cooking, households might use firewood for lighting and heating. Furthermore, easy accessibility to firewood in rural parts might make this attribute insignificant unlike in urban areas (Bielecki and Wingenbach 2014).

The coefficients of the scale heterogeneity multinomial logit model (SMNL) confirm the existence of scale heterogeneity (significant tau). Furthermore, the results of the generalized multinomial logit model show the existence of both scale and preference heterogeneity. Thus, accounting for these will improve the efficiency of the estimated coefficients.

The mixed logit model assumes continuous preference heterogeneity in estimating random parameters. This, however, may not always be true in the real world where people may exhibit discrete preference heterogeneity. Thus, this paper estimated the latent class model to allow a discrete change in preferences. Optimum class size was selected using the information criterion. Since the third and fourth classifications bring no improvements to the model’s fit, the two-class model was estimated. The results of this model are given in Table 10.

Finally, from the estimated models, the study computed marginal willingness to pay (MWT) for each attribute. This result carries important policy implications as cook stove designers, producers, and policymakers can clearly identify the most important features of the cook stove. The marginal rate of substitution (Table 11) shows the rate at which a consumer trades off one attribute for another.

As we can see in Table 11, the respondents made a trade-off when they took decisions about adopting the improved cook stove. They attached more value to the emission reduction attribute of the improved cook stove followed by reduction in the risks of a fire. Their marginal willingness to pay was lower for the fuel-saving attributes of the proposed cook stove.

This paper also examined the determinants of willingness to pay for the improved cook stove using data generated by the contingent valuation survey. For this, it analyzed the double-bounded contingent valuation data using the interval data model. For comparison purposes, we estimated the probit model using initial bid and the biprobit model using both responses to initial and follow-up bids and the interval data model. The results of the probit model are given in Table 12.

As expected, when the bid is increased, the probability of saying yes decreases. The educational level of a household head, household income, livestock ownership, and non-farm income positively affected a household’s willingness to pay for the improved cook stove.

The price of the stove (bid), family size, and off-farm employment negatively and significantly affected the adoption of the improved cook stove. This is similar to Mamuye et al.’s (2018) and Takema et al.’s (2011) findings. As expected, provision of cheap-talk reduced willingness to pay. This means that households who received the questionnaire with cheap-talk as an option stated lower willingness to pay compared to those without the cheap-talk option. Thus, cheap-talk can help reduce hypothetical bias in contingent valuations.

After estimating the determinants of willingness to pay for adopting the improved cook stove, this study also estimated mean willingness to pay from the different models estimated. The results of mean willingness to pay are presented in Table 13.

As we can see in Table 13, mean willingness to pay for the improved cook stove ranged between 152.66 birr and 353.87 birr. The actual market price of the improved cook stove in

### Table 7: Responses to distribution and the price of the stove (bid)

| Bid in ETB | No | Yes | Total |
|-----------|----|-----|-------|
| 75        | 26 | 79  | 105   |
| 150       | 50 | 50  | 100   |
| 250       | 62 | 40  | 102   |
| Total     | 138| 169 | 307   |

Source: Authors’ computations using survey data

### Table 8: Stated attribute non-attendance

| Attributes         | Attended to attribute (percentage of respondents) |
|--------------------|-----------------------------------------------|
|                    | Always | Often | Sometimes | Rarely | Never  |
| Emission reduction | 36.16  | 21.17 | 28.34     | 12.70  | 1.63   |
| Fuel wood saving   | 27.69  | 18.24 | 33.88     | 19.22  | 0.98   |
| Reducing risk of fire | 48.53 | 21.17 | 24.76     | 4.56   | 0.98   |
| Durability of the stove | 50.49 | 20.85 | 21.50     | 6.51   | 0.65   |

Source: Authors’ computations using survey data
Table 9  Results of the various discrete choice models

| Variables                          | Multinomial logit (MNL) | Mixed logit model (MXL) | Scale heterogeneity (SMNL) | Generalized multinomial logit (GMNL) |
|------------------------------------|-------------------------|-------------------------|---------------------------|--------------------------------------|
|                                   | Coefficient | SE       | Coefficient | SE       | Coefficient | SE       | Coefficient | SE       | Coefficient | SE       |
| ASC                                | 0.89***     | 0.28***  | 0.6  | 0.6  | 0.44***     | 0.4  | −0.04       | 0.12     | 1.50***     | 0.69     |
| Emission reduction                | 0.24***     | 0.06     | 0.12 | 0.12 | 0.27        | 0.09 | 0.27        | 0.02     | 1.27        | 0.23     |
| Fuel saving                       | 0.07        | 0.06     | 0.07 | 0.07 | 0.10        | 0.07 | 0.10        | 0.07     | 0.10        | 0.10     |
| Risk of fire                      | 0.84***     | 0.14     | 0.12 | 0.12 | 0.94***     | 0.09 | 0.94***     | 0.09     | 0.94***     | 0.17     |
| Durability of the stove           | 0.09***     | 0.01     | 0.09 | 0.09 | 0.01        | 0.01 | 0.01        | 0.01     | 0.01        | 0.01     |
| Price of the stove                | 0.002**     | 0.001    | 0.00 | 0.00 | 0.00        | 0.00 | 0.00        | 0.00     | 0.00        | 0.00     |
| SD emissions                      | 1.36***     | 0.14     | 0.17 | 0.17 | 0.003**     | 0.17 | 0.003**     | 0.17     | 0.003**     | 0.17     |
| SD fuel                           | 0.85***     | 0.14     | 0.19 | 0.19 | 0.001       | 0.19 | 0.001       | 0.19     | 0.001       | 0.19     |
| SD risk                           | 0.91***     | 0.01     | 0.14 | 0.14 | 0.001       | 0.14 | 0.001       | 0.14     | 0.001       | 0.14     |
| Tau (γ)                           | 0.06***     | 0.01     | 0.01 | 0.01 | 0.06***     | 0.01 | 0.06***     | 0.01     | 0.06***     | 0.01     |
| Gamma (ν)                         | 0.00       | 0.00     | 0.00 | 0.00 | 0.00       | 0.00 | 0.00       | 0.00     | 0.00       | 0.00     |
| Sigma (σ)                         | 1.00**     | 0.46     | 1.00 | 1.00 | 1.00**     | 0.46 | 1.00**     | 0.46     | 1.00**     | 0.46     |
| No. of parameters                 | 6          | 9        | 10   | 10   | 6          | 9    | 6          | 9        | 6          | 9        |
| Log likelihood                    | −1112.20    | −1110.20 | −1022.90 | −1022.90 | −1112.20    | −1110.20 | −1022.90 | −1022.90 | −1112.20    | −1110.20 |
| Pseudo R²                         | 0.20        | 0.45     | 0.45 | 0.45 | 0.20        | 0.45 | 0.45        | 0.45     | 0.20        | 0.45     |
| AIC                               | 228.60      | 229.60   | 229.60 | 229.60 | 228.60      | 229.60 | 229.60      | 229.60   | 228.60      | 229.60   |
| BIC                               | 2073.20     | 2073.20  | 2073.20 | 2073.20 | 2073.20     | 2073.20 | 2073.20     | 2073.20  | 2073.20     | 2073.20  |
| HQIC                              | 2248.70     | 2248.70  | 2248.70 | 2248.70 | 2248.70     | 2248.70 | 2248.70     | 2248.70  | 2248.70     | 2248.70  |

Note: ***, **, and * show significance at the 1%, 5%, and 10% level of significance.
Ethiopia is higher than the estimated mean willingness to pay. Lack of affordability of the improved cook stove might discourage its adoption and deprive households of its multiple benefits—health benefits due to reduction in indoor air pollution, cost reduction because of energy efficiency, and a reduction of environmental problems such as deforestation and CO2 emissions. It will be possible to harness the multiple benefits of the improved cook stove if the improved cook stove can be made available at a lower cost. This suggests that a dissemination intervention for the improved cook stove requires price support (subsidization policy) for it to be effective.

### Conclusion and policy recommendations

This study used two stated preference survey techniques for analyzing households’ willingness to pay and their preferences for the improved cook stove in Ethiopia. It used different discrete choice models for estimating the determinants of the use of the improved cook stove from data generated using the choice experiment survey. Households exhibited scale and preference heterogeneity when making choice decisions. Our study used the scale heterogeneous multinomial logit model (SMLM) and the generalized mixed logit model (GMLM) to account for this heterogeneity. Our results show the existence of scale and preference heterogeneity as seen in a significant

| Table 10 | Results of the latent class model |
|----------|----------------------------------|
| Attributes/class | Class 1 | Class 2 |
| ASC | -0.076 | -3.143 |
| Emission reduction | 0.245*** | -0.151 |
| Fuel wood saving | 0.100 | -0.983** |
| Reducing fire risks | 0.863*** | 0.643** |
| Durability of the cook stove | 0.085*** | 0.044 |
| Price of the cook stove | -0.001*** | -0.016*** |

| Class membership function: |
| Constant | 16.669*** |
| Income | 0.005*** |
| Age | -0.325*** |
| Latent class probs. | 0.961 | 0.039 |
| Number of obs. | 1842 |
| Log likelihood | -1074.22 |
| Pseudo R² | 0.29 |

Note: ***, **, and * show significance at the 1%, 5%, and 10% level of significance

| Table 11 | Marginal willingness to pay (MWT) for the improved cook stove’s attributes using the MNL model |
|-----------|------------------------------------------|
| Model/attributes | Emission reduction | Fuel saving | Risk reduction | Durability |
| MWT_MNL | 117.98*** | 36.49 | 472.58*** | 46.35*** |
| Std. dev. | (42.98) | (35.61) | (127.18) | (12.39) |

| Table 12 | Determinants of WTP’s results from CVM data |
|-----------|---------------------------------------------|
| Variables | Coef. | Std. err. |
| Stove price (bid) | -0.08*** | 0.002 |
| Age of the respondent (log) | 0.04 | 0.39 |
| Education level | 0.06* | 0.03 |
| Family size (AME) | -0.16** | 0.06 |
| Income (log) | 0.69*** | 0.20 |
| Livestock (TLU) | 0.18*** | 0.05 |
| Off-farm employment | -0.48* | 0.28 |
| Non-farm employment | 0.15* | 0.09 |
| Land owned | -0.03 | 0.02 |
| Access to credit | 0.26 | 0.21 |
| Own a private tree | 0.38 | 0.28 |
| Cheap-talk | -0.64*** | 0.21 |
| Constant | -2.95 | 2.11 |
| Pseudo R² | 0.29 |
| Likelihood ratio test, chi² test value | 85.20 (p-value 0.000) |

Note: ***, **, and * represent coefficients significant at the 1%, 5%, and 10% level respectively

| Table 13 | Results of mean willingness to pay using different models |
|-----------|----------------------------------|
| Variable/model | Probit | Probit with covariates | Doublelog | Doublelog model with covariate |
| Mean willingness to pay | 353.87*** | 208.56*** | 276.56*** | 152.66*** |
| Standard error | 47.72 | 44.75 | 10.45 | 27.72 |
tau coefficient. None of the existing studies on the topic have used these models and their results might be biased.

Our study also tested attribute non-attendance which is a behavioral response which could affect the estimated marginal rate of substitution among the different attributes. Using the stated attribute non-attendance, the study verified that households attended to all attributes when taking choice decisions. Fuel wood saving was less attended to among all the attributes. Our study also used cheap-talk to reduce the hypothetical bias from the contingent valuation exercise. Households who had received the questionnaire with cheap-talk as an option stated a lower mean willingness to pay compared to those without the cheap-talk option. Therefore, use of cheap-talk in contingent valuation will significantly reduce hypothetical bias.

Our results show that cook stove–related attributes and the respondents’ socioeconomic characteristics were important in determining preferences for adopting the improved cook stove. Emission reductions, lower risks of use, and durability of the stove positively affected the probability of households adopting the improved cook stove. Our results are similar to Mamuye et al.’s (2018) and Takema et al.’s (2011) findings. However, Bielecki’s and Wingenbach’s (2014) study stresses the importance of considering determinants other than socioeconomic and stove-related ones of the improved cook stove. Rural households value stoves beyond their cooking needs such as for lighting, heating, and as a gathering point for the family. Thus, stove designers need to consider these aspects too as they affect the adoption and dissemination of improved cook stoves. Among the attributes, the estimated marginal rates of substitution showed the trade-off that households made while deciding about purchasing the improved cook stove. Thus, cook stove designers, producers, and pricing policies should take this trade-off into account for effective dissemination of the improved stove. The estimated results of the scale heterogeneous multinomial logit model clearly show the existence of scale heterogeneity and the generalized multinomial logit model’s coefficients show that there are both scale and preference heterogeneities. Thus, failure to account for these factors will underestimate the coefficients of product-related and socioeconomic characteristics in estimating willingness to pay for the improved cook stove.

Our results also show that the sampled respondents were aware of the side effects of traditional energy sources and their health, environmental, and economic consequences. They were interested in adopting and using the improved cook stove but were frustrated by low-quality products and related inconveniences of using the limited existing products. Thus, designers of the improved cook stove should make it simple and convenient for use fitting the realities of rural housing and stove use. Reliability of the stove should also be ensured to enhance its uptake.

The estimated mean willingness to pay ranged from about 150 Birr to 350 Birr. This shows that the respondents’ willingness to pay was below the supply price of the improved cook stove in the area. Affordability of the improved cook stove might discourage its adoption and deprive households of its multiple benefits. Making the improved cook stove affordable for low-income rural households requires different price stabilizations. Capitalizing on the multiple benefits of the improved cook stove price subsidization can be obtained from sources such as carbon financing of the clean development mechanism.

Author contribution The authors’ individual contributions to the paper are as follows: Mekonnen Bersisa: Conceptualization, methodology, data curation, software, writing - original draft, writing - reviewing and editing, and project administration. Almas Heshmati: Conceptualization, methodology, supervision, writing - reviewing and editing, and validation. Alemu Mekonnen: Writing - reviewing and editing, and validation. All authors have approved the manuscript.

Funding Open access funding provided by Jönköping University.

Data availability The decoded data and codes are available on request.

Declarations

Ethics approval and consent to participate Considered

Consent for publication Not applicable

Competing interests The authors declare no competing interests.

Open Access This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons licence, and indicate if changes were made. The images or other third party material in this article are included in the article’s Creative Commons licence, unless indicated otherwise in a credit line to the material. If material is not included in the article’s Creative Commons licence and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this licence, visit http://creativecommons.org/licenses/by/4.0/.

References

Abdullah S, Jeanty PW (2011) Willingness to pay for renewable energy: evidence from a contingent valuation survey in Kenya. Renew Sust Energ Rev 15:2974–2983

Ackerman F, Heinzerling L (2002) Pricing the priceless: cost-benefit analysis of environmental protection. Univ Pa Law Rev 150(5):1553–1584

Adamowicz W, Boxall P, Williams M, Louviere J (1998) Stated preference approaches for measuring passive use values: choice experiments and contingent valuation. Am J Agric Econ 80(1):64–75
Agimas F, Mekonnen A (2011) Low-income fishermen’s willingness-to-pay for fisheries and watershed management: an application of choice experiment to Lake Tana, Ethiopia. Ecol Econ 71:162–170
Aizak H, Nishimura K (2008) Design and analysis of choice experiments using R: a brief introduction. Agricultural Information Research 17(2):86–94
Akpalu W, Dasmani I, Aglobitse PB (2011) Demand for cooking fuels in a developing country: to what extent do taste and preferences matter? Energy Policy 39:6525–6531
Alem Y, Hassen S, Köhlin G (2014) Adoption and dis-adoption of electric cook stove in urban Ethiopia: evidence from panel data. Resour Energy Econ 38:110–124
Amigun B, Musango JK, Stafford W (2011) Biofuels and sustainability in Africa. Renew Sust Energ Rev 15(2):1360–1372
Ben-Akiva M, Lerman S (1985) Discrete choice analysis: theory and application to travel demand. MIT Press, Cambridge
Bersisa M (2017) Multidimensional measure of household energy poverty and its determinants in Ethiopia. In: Heshmati A (ed) Economic Transformation for Poverty Reduction in Africa: A Multidimensional Approach. Routledge, New York, pp 58–83
Beyene, A., Bluffstone, R., Gebreegziaber, Z., Martinsson, P., Mekonnen, A., and Vieder, F. (2015). Do improved biomass cook stoves reduce fuelwood consumption and carbon emissions? Evidence from rural Ethiopia using a randomized treatment trial with electronic monitoring. World Bank, Policy Research Working Paper, WPS7324
Bhojvaid V, Jeuland M, Kar A, Lewis J, Patange O, Pattanayak S (2015) How do people in rural India perceive improved stoves and stoves reduce fuelwood consumption and carbon emissions? Evidence from rural Ethiopia using a randomized treatment trial with electronic monitoring. World Bank, Policy Research Working Paper, WPS7324
Bizzarri, M. (2010). Safe access to firewood and alternative energy in Ethiopia: an appraisal report. Prepared for the World Food Program. Available at: http://www.genderconsult.org/uploads/publications/doc/SAFE_Ethiopia_Appraisal_Report_Final_Draf2.pdf (Accessed on 10 March 2017)
Boxall PC, Adamowicz WL (2002) Understanding heterogeneous preferences in random utility models: a latent class approach. Environ Plan Manag 54(8):1061–1076
Brown BW, Walker MB (1989) The random utility hypothesis and inference in demand systems. Econometrica 57(4):815–829
Cameron AC, Trivedi PK (2010) Microeconometrics using Stata. College Station (TX), Stata Press, Revised Edition
Campbell D, Hensher DA, Scarpa R (2011) Non-attendance to attributes in environmental choice analysis: a latent class specification. J Environ Plan Manag 54(8):1061–1076
Carson RT (2012) Contingent valuation: a practical alternative when prices aren’t available. J Econ Perspect 26(4):27–42
Dawit D (2012) Assessment of biomass fuel resource potential and utilization in Ethiopia: sourcing strategies for renewable energies. International Journal of Renewable Energy Research 2(1):131–139
FDRE (2011a) Country strategy paper 2011-2015. FDRE, Addis Ababa
FDRE (2011b) Ethiopia’s climate-resilient green economy: green economy strategy. FDRE, Addis Ababa
Fiebig DG, Keane MP, Louviere J, Wasi N (2010) The generalized multinomial logit model: accounting for scale and coefficient heterogeneity. Mark Sci 29(3):393–421
Greene WH, Hensher DA (2003) A latent class model for discrete choice analysis: contrasts with mixed logit. Transp Res B 37:681–698
Haab TC, McConnell KE (2002) Valuing environmental and natural resources: the econometrics of non-market valuation. Edward Elgar Publishing Limited, Northampton, MA, USA
Hanemann WM (1994) Valuing the environment through contingent valuation. J Econ Perspect 8(4):19–43
Hanemann WM, Loomis J, Kanninen B (1991) Statistical efficiency of double-bounded dichotomous choice contingent valuation. Am J Agric Econ 73(4):1255–1263
Hanley N, Wright RE, Adamowicz V (1998) Using choice experiments to value the environment: design issues, current experience and future prospects. Environ Resour Econ 11(3–4):413–428
Hanley N, Mourato S, Wright RE (2001) Choice modeling approaches: superior alternatives for environmental valuation. J Econ Surv 15(3):435–462
Hausman J, McFadden D (1984) Specification tests for the multinomial logit model. Econometrica 52(5):1219–1240
Hensher DA, Greene WH (2003) The mixed logit model: the state of practice. Transportation 30:133–176
Hensher DA, Greene WH (2010) Non-attendance and dual processing of common-metric attributes in choice analysis: a latent class specification. Empir Econ 39:413–426
Hensher DA, Rose JM, Greene WH (2005) Applied choice analysis: a primer. Cambridge University Press, New York
Hensher DA, Rose JM, Greene WH (2012) Inferring attribute non-attendance from stated choice data: implications for willingness to pay estimates and a warning for stated choice design experiment. Transportation 39:235–245.ch
Hole AR, Kolstad IR (2012) Mixed logit estimation of willingness to pay distributions: a comparison of models in preference and WTP space using data from a health-related choice experiment. Empir Econ 42:445–469
Jagger P, Jumbe C (2016) Stoves or sugar? Willingness to adopt improved cook stoves in Malawi. Energy Policy 92:409–419
Jeuland M, Bhojvaid V, Kar A, Lewis J, Patange O, Pattanayak S (2015) Preferences for improved cook stoves: evidence from rural villages in north India. Energy Econ 52:287–298
Koozer, S.H. (2014). Clean cooking: the value of Clean Cook stoves in Malawi. Environ Res Public Health 11:1341
Kroon JD, Brouwer B, Van Beukering P (2014) The impact of the household decision environment on fuel choice behavior. Energy Econ 44:236–247
Lancaster KJ (1966) A new approach to consumer theory. J Polt Econ 74:132–157
List JA, Gallet C (2001) What experimental protocol influence disparities between actual and hypothetical stated values? Environ Resour Econ 20:241–254
Lopez-Feldman, A. (2012). Introduction to contingent valuation using Stata. Centro de Investigacion y Docencia Economicas (Cide’), MPRA Paper No. 41018, posted 4. September 2012 19:36 UTC. Available at: http://mpra.ub.uni-muenchen.de/41018/. Accessed on 30 April 2020
Macmurray, F., Lemme B, Woldeamanuel T (2018) Emissions and fuel use performance of two improved stoves and determinants of their adoption in Dodola, southeastern Ethiopia. Sustainable Environment Research 28:32–38
McFadden D (1974) Conditional logit analysis of qualitative choice behavior. In: In Zarembka, Frontiers in econometrics. Academic Press, New York
McFadden D (1986) The choice theory approach to market research. Mark Sci 5(4):275–297
Murphy JJ, Allen PG, Stevens TH, Weatherhead D (2005) A meta-analysis of hypothetical bias in stated preference valuation. Environ Resour Econ 30:313–325
Perman R, Ma Y, McGilvray J, Common M (2003) Natural resource and environmental economics, 3rd edn. Person Education Limited, London

Ruiz-Mercado I, Masera O, Zamora H, Smith KR (2011) Adoption and sustained use of improved cook stoves. Energy Policy 39:7557–7566

Scarpa R, Gilbride TJ, Campbell D, Hensher DA (2009) Modelling attribute non-attendance in choice experiments for rural landscape valuation. Eur Rev Agric Econ 36(2):151–174

Shen J (2009) Latent class model or mixed logit model? A comparison by transport mode choice data. Appl Econ 41(22):2915–2924

Takama, T., Lambe, F., Johnson, F.X., Arvidson, A., Atanassov, B., Debebe, M., Nilsson, L., Tella, P., and Tsephel, S. (2011). Will African consumers buy cleaner fuels and stoves? A household energy economic analysis model for the market introduction of bio-ethanol cooking stoves in Ethiopia, Tanzania, and Mozambique. Stockholm: Research Report, Stockholm Environmental Institute. Available at https://www.sei-international.org/mediamanager/documents/Publications/SEI-ResearchReport-Takama-Will African Consumers Buy Cleaner Fuels and Stoves-2011.pdf, (Accessed on 30 April 2020).

Train KE (1998) Recreation demand models with taste differences over people. Land Econ 74(2):230–239

Train K, Weeks M (2005) Discrete choice models in preference space and willingness-to-pay space. In: Alberini A, Scarpa R (eds) Application of simulation methods in environmental and resource economics. Kluwer Academic Publishers, Dordrecht, pp 1–16

UNDP and WHO. (2009). The energy access situation in developing countries: a review on the least developed countries and Sub-Saharan Africa. New York

Vigolo V, Sallaku R, Testa F (2018) Drivers and barriers to clean cooking: a systematic literature review from a consumer behavior perspective. Sustainability 10(11):4322, 1–4322, 25

WHO (2002) Addressing the links between indoor air pollution, household energy and human health. World Health Organization, Washington, DC

WHO (2007) Indoor air pollution: national burden of disease estimates. WHO, France

Publisher’s note Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.