Risk of death associated with the use of conventional versus atypical antipsychotic drugs among elderly patients

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ABSTRACT

Background: Public health advisories have warned that the use of atypical antipsychotic medications increases the risk of death among elderly patients. We assessed the short-term mortality in a population-based cohort of elderly people in British Columbia who were prescribed conventional and atypical antipsychotic medications.

Methods: We used linked health care utilization data of all BC residents to identify a cohort of people aged 65 years and older who began taking antipsychotic medications between January 1996 and December 2004 and were free of cancer. We compared the 180-day all-cause mortality between residents taking conventional antipsychotic medications and those taking atypical antipsychotic medications.

Results: Of 37,241 elderly people in the study cohort, 12,882 were prescribed a conventional antipsychotic medication and 24,359 an atypical formulation. Within the first 180 days of use, 1,822 patients (14.1%) in the conventional drug group died, compared with 2,337 (9.6%) in the atypical drug group (mortality ratio 1.47, 95% confidence interval [CI] 1.39–1.56). Multivariable adjustment resulted in a 180-day mortality ratio of 1.32 (1.23–1.42). In comparison with risperidone, haloperidol was associated with the greatest increase in mortality (mortality ratio 2.14, 95% CI 1.86–2.45) and loxapine the lowest (mortality ratio 1.29, 95% CI 1.19–1.40). The greatest increase in mortality occurred among people taking higher (above median) doses of conventional antipsychotic medications (mortality ratio 1.67, 95% CI 1.50–1.86) and during the first 40 days after the start of drug therapy (mortality ratio 1.60, 95% CI 1.42–1.80). Results were confirmed in propensity score analyses and instrumental variable estimation, minimizing residual confounding.

Interpretation: Among elderly patients, the risk of death associated with conventional antipsychotic medications is comparable to and possibly greater than the risk of death associated with atypical antipsychotic medications. Until further evidence is available, physicians should consider all antipsychotic medications to be equally risky in elderly patients.

Methods: We conducted a cohort study involving all British Columbia residents aged 65 years or more who filled a first-recorded (index) prescription for an oral antipsychotic medication between January 1996 and December 2004 and were free of cancer. We compared the 180-day all-cause mortality between residents taking conventional antipsychotic medications and those taking atypical antipsychotic medications. We also examined whether the risk of death differed by dose or duration of drug use and by dementia status and residence in a nursing home.

Antipsychotic medications are disproportionately used in elderly populations and have been prescribed to over a quarter of US Medicare beneficiaries in nursing homes. Reasons for their use include dementia, delirium, psychosis, agitation and affective disorders, but much use is outside approved indications. In addition, there have been rapid shifts away from first-generation conventional agents (e.g., chlorpromazine, haloperidol and loxapine) to more actively marketed second-generation atypical agents (e.g., clozapine, olanzapine, quetiapine and risperidone).

In a public health advisory issued on June 15, 2005, Health Canada warned that, compared with placebo, atypical antipsychotic medications increased the risk of death by 60% in a pooled analysis of 17 short-term randomized controlled trials involving elderly patients with dementia. Health Canada requested that “all manufacturers of these drugs include a warning and description of this risk in the safety information sheet for each drug.” The advisory did not extend to conventional antipsychotic medications, although the US Food and Drug Administration (FDA) noted that this is an important issue to study in the future. In the absence of data on the risk of death posed by conventional antipsychotic medications, there is mounting concern that clinicians may switch their elderly patients to these agents, particularly since their replacement by the newer drugs occurred so rapidly and recently. On the basis of extrapolations mainly from younger populations, some have suggested that the conventional formulations could, in theory, pose risks equal to or greater than those associated with the newer, atypical drugs in elderly populations. A cohort study involving US Medicare patients eligible for state-funded low-income pharmacy assistance programs showed a 37% increase in the 180-day mortality associated with the use of conventional antipsychotic medications compared with atypical ones. However, patients enrolled in state-funded pharmacy assistance programs are not representative of the general elderly population, since on average they have lower incomes and higher morbidity and mortality. We conducted a population-based cohort study involving all elderly people in British Columbia to compare the short-term mortality between those prescribed a conventional antipsychotic medication and those prescribed an atypical antipsychotic medication. We also examined whether the risk of death differed by dose or duration of drug use and by dementia status and residence in a nursing home.
Jan. 1, 1996, and Dec. 31, 2004. To ensure a uniform 1-year eligibility period before filling the index prescription, all study subjects were required to have used at least 1 medical service and have filled at least 1 prescription in the two 6-month intervals before the index date. They could not have used an antipsychotic medication in the year before the index date. We restricted the analysis to include only new users of antipsychotic medications to guard against selection bias among prevalent users from early symptom emergence, drug intolerance or treatment failure.15 Patients with a diagnosis of cancer at the index date were excluded to avoid residual confounding introduced by selective prescribing of conventional antipsychotic medications (chlorpromazine, haloperidol) as antiemetics in the most serious cancer cases, because these patients are also more likely to die within 180 days independent of drug use.

Patients were identified in linked administrative data from the BC Ministry of Health that contained information on all physician services (Medical Services Plan), hospital admissions with up to 25 diagnostic codes and all dispensings of prescription drugs, independent of payor, recorded by the province-wide PharmaNet database. We further linked vital status information from the BC Vital Statistics Agency, the provincial vital statistics bureau. Underreporting and misclassification of data appear to be minimal because of the electronic data entry of all drug dispensings and because hospital diagnoses showed good specificity and completeness.14 Linkage, performed with the use of a personal health number unique to every BC resident, is considered complete among patients using the provincial health care system. All traceable personal identifiers were removed to protect patient confidentiality. The Institutional Review Board of the Brigham and Women's Hospital approved this study, and data-use agreements with the BC Ministry of Health were in place.

Atypical antipsychotic agents17 included in the analyses were risperidone (74.7% of all atypical agents dispensed), quetiapine (14.9%), olanzapine (10.1%) and clozapine (0.3%). Other antipsychotic medications were considered to be conventional17 and included loxapine (69.4% of all conventional antipsychotic drugs dispensed), haloperidol (11.0%), chlorpromazine (7.4%), trifluoperazine (5.0%), thioridazine (3.1%), pimozide (2.4%), promazine (2.4%), perphenazine (1.5%), fluphenazine (0.2%), mesoridazine (0.1%) and thiothixene (< 0.1%). We converted daily doses to chlorpromazine-equivalent milligrams using the midpoints of recommended ranges in geriatric prescribing guidelines.18 We used the median daily dose in the population as a cutoff to assess the effect of higher and lower doses.

The study outcome was death from any cause, as recorded by the BC Vital Statistics Agency. A set of potential confounders was measured based on health care utilization data within 6 months before the initiation of index drug use (index date). These confounders included sociodemographic characteristics (age, sex, race, nursing home residence), generic markers of comorbidity that have shown good validity in predicting death recent work,27 as indicated by their most recent new prescription of antipsychotic drug). Using 2-stage linear regression for the instrumental variable estimation and additional adjustment for...
measured patient characteristics, we calculated the risk difference of 180-day mortality between patients using conventional and those using atypical antipsychotic medications. Linear regression to estimate risk difference is valid in large samples such as ours. Because patient-level observations were clustered in physician practices, we performed robust calculations of standard errors of the regression parameters to account for the within-physician correlation of outcomes (Stata statistical software, version 9, StataCorp LP, College Station, Tex.).

Results

A total of 37,241 elderly people in British Columbia were included in the study cohort; of these, 12,882 were prescribed a conventional antipsychotic medication and 24,359 an atypical formulation. The use of antipsychotic drugs increased during the study period, from 1.5 per 100 seniors to 2.5 per 100 seniors (Fig. 1). The use of atypical antipsychotic medications increased particularly rapidly and exceeded the use of conventional formulations in January 2000. Subjects in the conventional drug group were slightly younger and less likely to be male than those in the atypical drug group (Table 1). Those in the conventional drug group were less likely than those in the atypical drug group to have cerebrovascular disease, diabetes, acute MI, other cardiovascular diseases, dementia, delirium, psychoses, mood disorders and other psychiatric disorders; however, they were more likely to have congestive heart failure and non-MI ischemic heart disease at baseline. Compared with the atypical drug users, the conventional drug users had lower rates of antidepressant use and use of other psychotropic medications, lower total number of drugs, and fewer hospital admissions and nursing home stays.

Within the first 180 days of use of antipsychotic medications, the risk of death in the conventional drug group was 14.1%, compared with 9.6% in the atypical drug group (Table 2), for an unadjusted mortality ratio of 1.47 (95% confidence interval [CI] 1.39–1.56).

The adjusted mortality ratios comparing the risk of death between the conventional and atypical antipsychotic drug groups are shown in Table 3. Mortality within 180 days was meaningfully increased in the conventional drug group compared with the atypical drug group in multivariable models that controlled for a large number of potential confounders (mortality ratio 1.32). When comparing the most frequently prescribed antipsychotic drugs individually with risperidone, we found increased mortality ratios for haloperidol (2.14, 95% CI 1.86–2.45) and loxapine (1.29, 95% CI 1.19–1.40) but no difference for olanzapine (0.94, 95% CI 0.80–1.09). Yearly adjusted mortality ratios varied little, and in a nonsystematic way, from 1997 to 2004 (Fig. 2). The greatest increase in adjusted mortality risk for conventional versus atypical antipsychotic medications occurred with the use of higher (above median) conventional doses (mortality ratio 1.67) and during the first 40 days after start of therapy (mortality ratio 1.60). In analyses restricted by dementia status or residence in a nursing home, patients in the conventional antipsychotic drug group consistently had a greater 180-day mortality than did patients in the atypical drug group (Table 3). A multivariable analysis of the difference in mortality revealed an estimated increase of 3.5 deaths per 100 population (95% CI 2.7–4.3) in the conventional drug group.

Our confirmatory analyses using propensity score adjustments yielded no substantive changes relative to the traditional multivariable Cox regression analyses. For example, the mortality ratio comparing the risk of death within 180 days between the conventional and atypical antipsychotic drug groups after propensity score adjustment was 1.39 (95% CI 1.30–1.49).

In the instrumental variable analyses, use of conventional antipsychotic medications continued to be associated with an increased risk of death within 180 days compared with use of atypical formulations. In these analyses, the adjusted risk difference of 4.2 per 100 (95% CI 1.2–7.3) meant that, for every 100 patients prescribed a conventional antipsychotic drug instead of an atypical drug, there were about 4 additional deaths. The adjusted estimates in the instrumental variable analyses did not differ from the traditional multivariable estimates (p = 0.62). Our instrument was strongly associated with the actual treatment choice (odds ratio 6.1, 95% CI 5.8–6.4).

Sensitivity analyses to rule out possible bias from residual confounding revealed that very large relative risks of 5 or more linking a hypothetical confounder to both conventional antipsychotic drug use and mortality would be needed to fully explain the observed increased mortality associated with conventional antipsychotic drug use if no such increase existed.
Interpretation

In our study involving 37,241 elderly residents in British Columbia who began taking antipsychotic therapy during the study period, patients prescribed a conventional agent had a 32% greater, dose-dependent risk of death within 180 days than did those given an atypical agent. To place this magnitude of risk in perspective, all measured health conditions except congestive heart failure and HIV infection conferred smaller adjusted mortality rate ratios in our analyses.

| Table 1: Characteristics of 37,241 elderly residents of British Columbia who began taking antipsychotic medications between Jan. 1, 1996, and Dec. 31, 2004 |
|---------------------------------|-----------------|-----------------|-----------------|-----------------|
| Characteristic                 | Type of antipsychotic medication; no. (%) of patients* | Conventional n = 12,882 | Atypical n = 24,359 | Odds ratio | p value |
| Age, yr, mean (SD)             |                 | 79.9 (8.8)      | 80.3 (8.4)      | 1.47 (1.39–1.56) | < 0.001 |
| Sex, male                      |                 | 5120 (39.7)     | 8565 (35.2)     | 0.9 (0.84–1.56)  | < 0.001 |
| History                        |                 |                 |                 | 0.5 (0.4–0.6)    | < 0.001 |
| Cardiac arrhythmia             |                 | 6 (0.0)         | 22 (0.1)        | 0.4 (0.3–0.6)    | 0.14 |
| Cerebrovascular disease        |                 | 1391 (10.8)     | 2430 (10.0)     | 0.9 (0.84–1.06)  | 0.013 |
| Congestive heart failure       |                 | 1084 (8.4)      | 1455 (6.0)      | 1.2 (1.1–1.4)    | < 0.001 |
| Diabetes mellitus              |                 | 1939 (15.1)     | 3362 (13.8)     | 0.9 (0.84–1.06)  | 0.001 |
| Myocardial infarction          |                 | 354 (2.7)       | 551 (2.3)       | 0.2 (0.1–0.3)    | 0.052 |
| Other ischemic heart disease   |                 | 493 (3.8)       | 665 (2.7)       | 1.1 (1.0–1.2)    | < 0.001 |
| Other cardiovascular disorder  |                 | 2609 (20.3)     | 4075 (16.7)     | 1.0 (0.94–1.06)  | < 0.001 |
| Cancer                         |                 | 0               | 0               | —                | —     |
| HIV infection                  |                 | 0               | 4 (0.0)         | 0.0 (0.0–0.1)    | 0.15 |
| Dementia                       |                 | 1247 (9.7)      | 3087 (12.7)     | 0.6 (0.55–0.7)   | < 0.001 |
| Delirium                       |                 | 967 (7.5)       | 2060 (8.5)      | 0.7 (0.6–0.8)    | < 0.002 |
| Mood disorder                  |                 | 2013 (15.6)     | 6198 (25.4)     | 0.4 (0.36–0.5)   | < 0.001 |
| Psychotic disorder             |                 | 1446 (11.2)     | 4103 (16.8)     | 0.5 (0.44–0.6)   | < 0.001 |
| Other psychiatric disorder     |                 | 403 (3.1)       | 1110 (4.6)      | 0.6 (0.54–0.7)   | < 0.001 |
| Use of other drugs             |                 |                 |                 |                 |       |
| Antidepressant                 |                 | 3645 (28.3)     | 10154 (41.7)    | 0.4 (0.36–0.49)  | < 0.001 |
| Other psychotropic medication  |                 | 542 (4.2)       | 920 (3.8)       | 0.9 (0.84–0.96)  | 0.040 |
| Prior anticholinergic drug use  |                 | 1140 (8.9)      | 1709 (7.0)      | 0.8 (0.72–0.9)   | < 0.001 |
| Current anticholinergic drug use|                 | 1868 (14.5)     | 2591 (10.6)     | 0.7 (0.62–0.79)  | < 0.001 |
| Total no. of drugs used, mean (SD) |             | 7.4 (5.1)       | 7.3 (5.0)       | —                | 0.60 |
| Hospital admission in previous 180 d |             | 1923 (14.9)     | 3204 (13.2)     | 0.9 (0.84–1.06)  | < 0.001 |
| Nursing home residence in previous 180 d |       | 3980 (30.9)     | 6471 (26.6)     | 0.9 (0.84–1.06)  | < 0.001 |

*Unless stated otherwise.

| Table 2: Mortality within 180 days after start of antipsychotic therapy |
|-------------------|-------------------|-------------------|-------------------|
| Variable           | Conventional n = 12,882 | Atypical n = 24,359 | Ratio (95% CI) | Difference (95% CI) |
| No. of patient-years | 5,816.4          | 11,354.3          | —                | —                |
| No. of deaths      | 1,822             | 2,337             | —                | —                |
| Risk of death per 100 population | 14.1 | 9.6 | 1.47 (1.39–1.56) | 4.55 (3.84–5.26) |
| Rate of death per 100 patient-years | 31.3 | 20.6 | 1.52 (1.43–1.62) | 10.70 (9.07–12.40) |

Note: CI = confidence interval.
Our results are remarkably close to the increased 180-day mortality associated with conventional antipsychotic drug use observed among US Medicare patients eligible for state-funded low-income pharmacy assistance programs (risk ratio 1.37, 95% CI 1.27–1.49). This association was confirmed shortly afterward in a meta-analysis of randomized trials. Compared with placebo, the conventional agent haloperidol increased the risk of short-term death by 107%—an estimate higher than that for atypical antipsychotic drugs and remarkably close to the increased risk of 60%–70% associated with atypical drugs versus placebo plus the 35% increase in risk associated with conventional drugs observed in our study.

Nonrandomized studies using health care utilization data are particularly scrutinized for their limited control of confounders and their potential for misclassifying diagnoses. Confounding would occur if patients who were frail and at increased risk of death were more likely to be prescribed a conventional antipsychotic medication than an atypical formulation. We therefore controlled for calendar year and sociodemographic, clinical and health care utilization factors likely to be independent predictors of death using traditional multivariable analyses as well as propensity score and instrumental variable analysis techniques. Our ability to adjust fully for those factors was limited by their measurement in our database. Random misclassification of confounders leads to incomplete adjustment of confounding bias. Model predictions of mortality based on measured covariates in the atypical and conventional antipsychotic user groups indicated non-differential assessment of patient characteristics. Despite our attempts to control for confounding by indication, it is possible that physicians prescribed haloperidol more frequently than a newer, atypical antipsychotic medication for acute agitation in patients who were more likely to die. It is also possible that patients using atypical antipsychotic medications were physically and cognitively more impaired.

We restricted our study population to new users of antipsychotic medications to control for indications and to ensure that the chronology of use was aligned in both groups and that patient characteristics were measured before drug use and thus not influenced by any treatment effects. We further analyzed data as intention-to-treat because of the known potential for drug intolerance or treatment failure that may lead to informative censoring. Such intention-to-treat analyses will likely bias results toward the null. During the study period (i.e., before the FDA health advisory was posted in 2005), recommendations were published to avoid the prescription of conventional antipsychotic medications to frail elderly patients, and any residual confounding may have therefore led to an underestimation of mortality associated with the use of conventional agents.

Finally, we used instrumental variable analysis, which by design can control for unmeasured patient characteristics, and which confirmed our results. Like other statistical approaches, the validity of instrumental variable estimation relies on assumptions. First, the instrument must be related to the actual exposure, which we could demonstrate in our study. Second, an instrument must not be correlated with patient risk factors conditional on measured and adjusted covariates. We found that large imbalances of risk factors among the actual treatment groups (Table 1) were substantially reduced in the instrumental variable analysis (data not shown). Although we have shown earlier how methods for instrumental variable estima-

| Model                                                 | Mortality ratio (95% CI) |
|-------------------------------------------------------|--------------------------|
| Unadjusted analysis                                    | 1.47 (1.39–1.56)         |
| Adjusted for age, sex and calendar year                | 1.11 (1.04–1.19)         |
| Multivariable analysis*                                |                          |
| Use of any conventional APM                            | 1.32 (1.23–1.42)         |
| Use of high-dose conventional APM                      | 1.67 (1.50–1.86)         |
| Use of low-dose conventional APM                       | 1.23 (1.14–1.33)         |
| Multivariable analysis by duration of use*             |                          |
| < 40 d after start of therapy                          | 1.60 (1.42–1.80)         |
| 40–79 d after start of therapy                         | 1.31 (1.14–1.51)         |
| 80–180 d after start of therapy                         | 1.18 (1.06–1.31)         |
| Multivariable analysis by patient subgroup*            |                          |
| With dementia                                          | 1.26 (1.01–1.56)         |
| Without dementia                                       | 1.30 (1.21–1.40)         |
| In a nursing home                                      | 1.25 (1.12–1.40)         |
| Not in a nursing home                                  | 1.35 (1.23–1.49)         |

Note: CI = confidence interval, APM = antipsychotic medication.

*Hazard ratios were adjusted for calendar year, age, sex, race, presence or absence of cardiac arrhythmias, cerebrovascular disease, congestive heart failure, diabetes, myocardial infarction, other ischemic heart disease, other cardiovascular disorders, cancer, HIV infection, dementia, delirium, mood disorders, psychotic disorders, other psychiatric disorders, and the use or nonuse of other psychiatric medications, prior use of anticholinergic drugs, current use of anticholinergic drugs, total number of medications used, hospital admissions and nursing home stays.

Fig. 2: Yearly adjusted mortality ratios comparing the risk of death between the conventional and atypical antipsychotic drug groups, from 1997 to 2004. Error bars represent 95% confidence intervals.
tion perform when using health care utilization databases to study the safety of prescription drug use,34 this does not rule out that some residual confounding persisted.

Non-differential exposure misclassification (e.g., not consuming filled prescriptions or switching classes of antipsychotic medication) and any rare misclassification of British Columbia mortality information would bias results toward the null; differential misclassification (e.g., worse adherence with conventional antipsychotic medication, as has been found35) again may lead to an underestimation of mortality associated with conventional agents. An alternative interpretation untestable in our data is that health care providers may have managed the indication using harsher co-interventions (e.g., physical restraint, sedatives) when conventional therapy failed.

Potential mechanisms through which conventional antipsychotic medications might increase short-term mortality are unknown. In the FDA analysis on which its public health advisory was based, heart-related events (heart failure, sudden death) and infections (mostly pneumonia) accounted for most deaths.6 Anticholinergic properties (affecting blood pressure and heart rate), Q–T prolongation (causing conduction delays) and extrapyramidal symptoms (causing swallowing problems) are at least as common and probably more common with conventional than with atypical antipsychotic agents and should be investigated as potential underlying causes.31-33

Together with earlier findings, the results from our study strongly suggest that Health Canada and the FDA should include conventional antipsychotic medications in their public health advisories, which currently warn only of the increased risk of death associated with the use of atypical antipsychotic medications in elderly patients with dementia.

This article has been peer reviewed.

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Competing interests: None declared.

Contributors: All authors substantially contributed to the conception, design, analysis and interpretation of data, drafting and revising the article for important intellectual content, and gave approval of the final version.

Acknowledgements: This project was funded by the DEcIDE research network by a contract from the Agency for Health Care Quality and Research to the Brigham and Women’s Hospital DEcIDE Research Center (HHS2 29200205016F). Sebastian Schneweiss was further supported by a grant from the National Institute on Aging (RO1-AG021950).

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