Endometriosis is a common gynecological condition with complex etiology defined by the presence of endometrial glands and stroma outside the womb. Endometriosis is a common cause of both cyclic and chronic pelvic pain, reduced fertility, and reduced quality-of-life. Diagnosis and treatment of endometriosis is, on average, delayed by 7–10 years from the onset of symptoms. Absence of a timely and non-invasive diagnostic tool is presently the greatest barrier to the identification and treatment of endometriosis. Twin and family studies have documented an increased relative risk in families. To identify genetic factors that contribute to endometriosis we conducted a two-stage genome-wide association study (GWAS) of a European cohort including 2,019 surgically confirmed endometriosis cases and 14,471 controls. Three of the SNPs we identify associated at \( P < 5 \times 10^{-8} \) in our combined analysis belong to two loci: LINC00339-WNT4 on 1p36.12 (rs2235529; \( P = 8.65 \times 10^{-9}, \text{OR} = 1.29, \text{CI} = 1.18–1.40 \)) and RND3-RBM43 on 2q23.3 (rs1519761; \( P = 4.70 \times 10^{-8}, \text{OR} = 1.20, \text{CI} = 1.13–1.29 \)). Using an adjusted Bonferroni significance threshold of 4.51 \( \times 10^{-5} \) we identify two additional loci in our meta-analysis that associate with endometriosis: RNF144B-ID4 on 6p22.3 (rs6907340; \( P = 2.19 \times 10^{-7}, \text{OR} = 1.20, \text{CI} = 1.12–1.28 \)), and HNRNPA3P1-LOC100130539 on 10q11.21 (rs10508881; \( P = 4.08 \times 10^{-7}, \text{OR} = 1.19, \text{CI} = 1.11–1.27 \)). Consistent with previously suggested associations to WNT4 our study implicates a 150 kb region around WNT4 that also include LINC00339 and CDC42. A univariate analysis of documented infertility, age at menarche, and family history did not show allelic association with these SNP markers. Clinical data from patients in our study reveal an average delay in diagnosis of 8.4 years and confirm a strong correlation between endometriosis severity and infertility (\( n = 1182, P = 0.001, \text{OR} = 2.18 \)). This GWAS of endometriosis was conducted with high diagnostic certainty in cases, and with stringent handling of population substructure. Our findings broaden the understanding of the genetic factors that play a role in endometriosis.

Introduction
Endometriosis affects 5–10% of women in their reproductive years with symptoms including pelvic pain, dyspareunia, dysmenorrhea and infertility [1]. Although ectopic endometrium has been observed in female fetuses [2], symptoms of endometriosis usually don’t manifest until adolescence, and some patients with severe endometriosis remain asymptomatic. Definitive diagnosis is often delayed 7–10 years after the onset of symptoms severely impacting quality of life [3–5]. Family history of endometriosis has been reported in multiple studies to increased relative risk about a 5-fold [6,7]. A large twin-study based on the Australian Twin Registry has shown that the ratio of mono-zygotic to fraternal twin pair correlations was in excess of 2 fold, suggesting that 51% of the variance of the liability to endometriosis may be attributable to additive genetic influences with minimal influence from environmental factors [8], and two smaller twin-studies report the concordance rate of endometriosis between monozygotic twins to range between 75% and 87% [9,10]. A large number of candidate genes have been investigated for their role in endometriosis as summarized by Montgomery et al. [11] and Rahmioglu et al. [12], but the first strong evidence to date for genetic association are reported in two large Genome-Wide Association Studies (GWAS). In the first study Uno et al. [13] identified rs10965235 located in an intron of CDKN2BAS on chromosome 9p21 to be associated in a Japanese cohort, and in the second study Painter et al. [14] identified the intergenic SNP rs12700667 on chromosome 7p15.2 to be associated in a European cohort. A meta-analysis of the two studies extend this evidence and identify a total of seven loci associated with endometriosis [15]. To replicate and extend our understanding of the genetic factors that contribute to endometriosis we have undertaken a large two-stage GWAS in a European cohort.

Results and Discussion
GWAS and Replication
We conducted a discovery GWAS on surgically confirmed endometriosis patients and population controls using the Illumina OmniExpress BeadChip. SNPs were limited to the autosomes and...
SNPs with an Illumina GenType score ≥0.65. We further eliminated SNPs with call rate <0.98, Hardy-Weinberg Equilibrium (hwe) <0.001 and minor allele frequencies (MAF) <0.01. After filtering 580,699 SNPs remained. Next, samples with identity-by-state (IBS), and samples closer than 3rd-degree (tested for unknown familial relationships using genome-wide Pmeta = 4.08) threshold including: LINC00339-WNT4 on 1p36.12 (rs2235529; Pmeta = 3.05) and 6p22.3 (rs6757804; Pmeta = 2.19) were removed. We used ADMIXTURE (ver. 1.22) [16] to estimate individual ancestry proportions based on a subset of SNPs on the Illumina OmniExpress chip (see Materials and Methods) and restricted our analysis to samples with ≥95% European ancestry. The calculated ancestral distribution of samples within Europe is shown in Figure S1. After applying quality, relatedness and ethnicity filters, 1,514 case and 12,660 control samples were used for the discovery phase of the association analysis. The genomic inflation factor lambda (λ) was determined to be 1.18, indicating measurable population stratification across the samples. To account for the elevated λ, we performed a PCA adjusted association analysis that resulted in a λ-value of 1.05 shown in QQ-plots in Figure S2. We selected the top 100 SNPs with the lowest PCA-adjusted P-values (ranging between 8.20×10^-5 and 1.36×10^-7) for further association analysis in the replication stage (Table S1).

The replication samples included 505 cases and 1811 controls selected for the same criteria as the discovery set. The λ-value for the replication cohort was determined to be 1.01 suggesting no measurable population stratification. After applying the same SNP filters as above we analyzed the top 100 SNPs from the discovery GWAS in the replication set. A significance threshold for the study, allowing for multiple correction, was chosen at 4.51×10^-7 (0.05/108,699; 108,699 being the number of independent SNPs in the panel of 580,699 filtered SNPs with r^2<0.20). A meta-analysis of the discovery and replication results was performed using Cochran-Mantel-Hanzel test and revealed 8 SNPs from 4 genomic regions that passed our genome-wide significance threshold including: LINC00339-WNT4 on 1p36.12 (rs2235529; Pmeta = 3.05×10^-9, OR = 1.30); RND3-RBM43 on 2q23.3 (rs6757804; Pmeta = 6.45×10^-8, OR = 1.20); RNFL14B-ID4 on 6p22.3 (rs6907340; Pmeta = 2.19×10^-7, OR = 1.20); and HNRNPA3P1-LOC100135039 on 10q11.21 (rs10508881; Pmeta = 4.08×10^-7, OR = 1.19) and shown in detail in Table 1. Table 1 also shows that three SNPs (rs2235529, rs15917561 and rs6757804) pass a conventional genome wide significance threshold of P<5×10^-8 in the combined analysis. There was no evidence of heterogeneity between the discovery and replication datasets as judged by the Breslow Day test and shown in Table 1. A second group of 15 SNPs from nine genomic regions that show suggestive replication provide additional candidate loci for endometriosis that merit further investigation. The most significant of these regions encompass IL33 on 9p24.1 (rs10975519; Pmeta = 9.26×10^-7, OR = 1.19). IL33 is a chemokine that has been linked to deep infiltrating endometriosis [17]. A summary of the discovery and replication GWAS, together with the meta and combined analysis for all 100 SNPs is presented in Table S1.

To further characterize the signals from the four most strongly associated regions and the IL33 region, we performed imputation using 1000-Genome and dataset utilizing IMPUTE2 (ver. 2.2.2) [18]. The results from the imputed dataset for the WNT4 region is shown in Table S2a. Figure 1 show a regional association plot around WNT4 and reveal that the associated region extends across 150 kb and include two new genes, HSPC157 (gene symbol: LINC00339) and CDC42, in addition to WNT4. The imputation results identify 25 additional SNPs that are strongly associated to endometriosis and reveal that the imputed SNP, rs1097151, is the most strongly associated SNP in the region (Pallele = 5.63×10^-10, OR = 1.3). WNT4 is important for steroidogenesis, ovarian follicle development and the development of the female reproductive tract and a very plausible candidate for endometriosis based on its biological functions [19]. Cell division cycle 42, CDC42, is a small GTPase of the Rho-subfamily, which regulates signaling pathways that control diverse cellular functions including cell morphology, migration, endocytosis and cell cycle progression. CDC42 is in part regulated by estrogen, expressed in endometrium, and has been shown to be differentially expressed in endometriosis [20]. HSPC157 is an alias for the Long Intergenic Non-protein Coding RNA 339 (gene symbol: LINC00339). HSPC157 has been found to be differentially expressed in endometriosis versus autologous uterine endometrium [21]. Based on this biological evidence both CDC42 and LINC00339 must also be considered candidates for endometriosis.

WNT4 has previously been associated with endometriosis, first by Uno et al. (2010) who noted that rs16826658, approximately 16 kb upstream of WNT4, showed a possible association to endometriosis (P = 1.66×10^-6, OR = 1.20) in a Japanese population, and again by Painter et al. (2011) who reported that rs7210902, located approximately 22 kb upstream of WNT4, also showed evidence for association in a European cohort (P = 9.0×10^-5, OR = 1.16). Our imputation analysis replicate the association of rs7210902 with endometriosis (P = 6.4×10^-5, OR = 1.17) and confirm the involvement of the WNT4-region in the pathogenesis of endometriosis. We only found weak evidence of association with rs16826658 (P = 0.05, OR = 1.07), because the minor allele is very common and because of the different ethnic backgrounds between the studies. To further evaluate the signals from the WNT4 region, we performed a haplotype analysis of three key SNPs from our study (rs10917151, rs4654783 and rs2235529) together with rs16826658, and rs7210902 using the imputed data from our population. The haplotype-results are summarized in Table S3 and show that the risk for endometriosis is confined to a single haplotype anchored in rs10917151, rs4654783 and rs2235529 (PHAP-1 = 7.26E-09, ORHAP-1 = 1.28), and that this haplotype starts to deteriorate with the addition of rs16826658 and rs7210902 (PHAP-1 = 8.33E-07, ORHAP-1 = 1.25). This analysis suggests that the risk allele observed in the present study and the two previously published reports is located on the same ancestral haplotype. Imputation was also performed on the three other significant regions on chromosome 2, 6, and 10, and on the region surrounding IL33 on chromosome 9 (Table S2a–e, Figure S3).

Overlap with Other Reported Loci
Uno et al. [13] reported association between endometriosis and rs10965235, located in intron 19 of CDKN2BAS, in a Japanese population. The protective minor allele A, that has a minor allele frequency of 0.20 in the Japanese population, is not observed in the European population preventing a direct comparison between the two ethnic groups. In lieu of a direct comparison between the two ethnic groups we scanned 66 SNPs from a region of 200 kb surrounding rs10965235 in our European population. After correcting for multiple-testing (P-value ≤0.05/66 = 0.00076) we didn’t find any evidence that CDKN2BAS is associated with endometriosis in the European population (Table S4a). In contrast, a second SNP (rs13271465) located on 8p22 between MTMR7 and SLC7A2, that Uno et al. identified as being associated with endometriosis (Pcombined = 9.84×10^-6, OR = 1.18), is present in both studies, show weak association (P = 0.0057, OR = 1.14) in our study, while none of the other 88 SNPs from the 200 kb region surrounding rs13271465 showed
any evidence for association with endometriosis (Table S4b). Uno et al. provided a supplementary list of 100 additional candidate SNPs from their study of which 60 are present in our analysis. Among the 60 SNPs only rs2473277 has a P-value $<0.001$ in our study ($P = 5.97 \times 10^{-6}$, OR = 1.17). SNP rs2473277 is located between LINC00339 and CDC42 at the left-most boundary of the WNT4 LD-block discussed above (Figure 1), and the risk allele of rs2473277 tag perfectly with the risk haplotype HAP-1 shown in Table S3 (data not shown).

Painter et al. [14] reported significant association of moderate and severe endometriosis with rs12700667 on 7p15.2 ($P_{adj} = 2.6 \times 10^{-7}$, OR = 1.22), with flanking support from

Table 1. Summary of GWAS and replication results.

| SNP       | gene          | Chr | Pos       | allele | stage   | Case MAF | Control MAF $^a$ | OR $^b$ | 95%CI    | $P_{het}$ $^d$ |
|-----------|---------------|-----|-----------|--------|---------|----------|-----------------|--------|---------|---------------|
| rs4654783 | WNT4          | 1   | 22,439,520| a/g    | Discovery | 0.34     | 0.295          | 2.43E-07| 1.23     | 1.14–1.34     |
|           |               |     |           |        | Replication | 0.323   | 0.298          | 1.31E-01| 1.13     | 0.97–1.32     |
|           |               |     |           |        | Meta $^a$   | 1.40E-07| 1.21          | 1.13–1.30| 0.332   |               |
|           |               |     |           |        | Combined Trend $^d$ | 0.336 | 0.295          | 1.17E-07| 1.21     | 1.13–1.29     |
| rs2235529 | WNT4          | 1   | 22,450,487| a/g    | Discovery | 0.188    | 0.153          | 1.36E-07| 1.28     | 1.16–1.41     |
|           |               |     |           |        | Replication | 0.182   | 0.142          | 1.38E-03| 1.36     | 1.13–1.64     |
|           |               |     |           |        | Meta $^a$   | 3.05E-09| 1.3            | 1.19–1.41| 0.583   |               |
|           |               |     |           |        | Combined Trend $^d$ | 0.186 | 0.151          | 8.65E-09| 1.29     | 1.18–1.40     |
| rs1519754 | RND3 | RBM43 | 2   | 151,619,693| c/a   | Discovery | 0.446    | 0.403          | 5.67E-05| 1.19     | 1.11–1.30     |
|           |               |     |           |        | Replication | 0.45    | 0.405          | 9.63E-03| 1.2      | 1.04–1.38     |
|           |               |     |           |        | Meta $^a$   | 1.75E-07| 1.2            | 1.12–1.28| 0.964   |               |
|           |               |     |           |        | Combined Trend $^d$ | 0.447 | 0.403          | 1.15E-07| 1.2      | 1.12–1.28     |
| rs6734792 | RND3 | RBM43 | 2   | 151,624,882| g/a   | Discovery | 0.448    | 0.404          | 3.52E-05| 1.2      | 1.11–1.29     |
|           |               |     |           |        | Replication | 0.453   | 0.406          | 7.48E-03| 1.21     | 1.05–1.39     |
|           |               |     |           |        | Meta $^a$   | 8.18E-08| 1.2            | 1.12–1.28| 0.945   |               |
|           |               |     |           |        | Combined Trend $^d$ | 0.449 | 0.404          | 5.19E-08| 1.2      | 1.12–1.28     |
| rs1519761 | RND3 | RBM43 | 2   | 151,633,204| g/a   | Discovery | 0.445    | 0.401          | 3.54E-05| 1.2      | 1.11–1.29     |
|           |               |     |           |        | Replication | 0.452   | 0.403          | 5.85E-03| 1.21     | 1.05–1.40     |
|           |               |     |           |        | Meta $^a$   | 7.30E-08| 1.2            | 1.12–1.29| 0.886   |               |
|           |               |     |           |        | Combined Trend $^d$ | 0.447 | 0.401          | 4.70E-08| 1.2      | 1.13–1.29     |
| rs6757804 | RND3 | RBM43 | 2   | 151,635,832| g/a   | Discovery | 0.445    | 0.401          | 3.43E-05| 1.2      | 1.11–1.29     |
|           |               |     |           |        | Replication | 0.452   | 0.403          | 5.44E-03| 1.21     | 1.06–1.40     |
|           |               |     |           |        | Meta $^a$   | 6.45E-08| 1.2            | 1.13–1.29| 0.876   |               |
|           |               |     |           |        | Combined Trend $^d$ | 0.446 | 0.4011         | 4.05E-08| 1.2      | 1.13–1.29     |
| rs6907340 | RNF144B | ID4 | 6   | 19,803,768| a/g   | Discovery | 0.417    | 0.371          | 5.49E-06| 1.21     | 1.12–1.31     |
|           |               |     |           |        | Replication | 0.412   | 0.378          | 4.54E-02| 1.15     | 1.00–1.33     |
|           |               |     |           |        | Meta $^a$   | 2.19E-07| 1.2            | 1.12–1.28| 0.579   |               |
|           |               |     |           |        | Combined Trend $^d$ | 0.415 | 0.372          | 1.25E-07| 1.2      | 1.12–1.28     |
| rs10508881| HNRNPA3P1 | LOC100130539 | 10 | 44,541,565| a/g   | Discovery | 0.45    | 0.405          | 3.18E-05| 1.2      | 1.11–1.30     |
|           |               |     |           |        | Replication | 0.42    | 0.387          | 6.06E-02| 1.15     | 1.00–1.32     |
|           |               |     |           |        | Meta $^a$   | 4.08E-07| 1.19            | 1.11–1.27| 0.589   |               |
|           |               |     |           |        | Combined Trend $^d$ | 0.442 | 0.403          | 1.57E-06| 1.18     | 1.10–1.26     |

The discovery stage included 1,514 endometriosis cases and 12,660 population controls, and the replication stage included 505 cases and 1,811 controls.

$^a$The P-values were determined using the Cochrane-Armitage trend test. P-values for the Discovery set reflect PCA adjusted P trend values.

$^b$Odds-ratios (OR) and confidence intervals (CI) are calculated using the non-risk allele as the reference.

$^c$The Meta analysis was performed using Cochran-Mantel-Haenzel statistics.

$^d$P values of heterogeneties ($P_{het}$) across discovery and replication stages calculated using Breslow-Day Test.

Significance threshold is $4.59 \times 10^{-7}$, and is determined by 0.05/108,699 where 108,699 is the number of independent SNPs in the panel with $r^2$ less than 0.20.

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rs7798431 located 41 kb away. The two SNPs were reported to be in strong linkage disequilibrium (LD), but unfortunately neither SNP is included in our study. A review of the Hapmap3 data from the region show that three SNPs in our study (rs12535837, rs10282436, rs10232819) are in moderate to strong LD with rs12700667 and rs7798431, but we find no evidence for association between endometriosis and any of these markers in our analyses of all endometriosis cases together nor do we find any evidence to the moderate and severe subset (Table S4c). A broader scan of the 200 kb region surrounding rs12700667, suggests weak association with rs4722551 (P = 0.000867) about 90 kb downstream of rs12700667 (Table S4c). Painter et al. provided a supplementary list of 73 candidate SNPs, but none of the thirty-nine SNPs present in our study reach a P-value threshold <0.001.

A recent meta-analysis published by Nyholt et al. [15] extend the findings by Uno et al. and Painter et al. and report a total of seven SNPs that pass a genome-wide significance (P<5x10^-8). A comparison of our results to each of the seven loci show evidence for association to endometriosis for three of the seven loci as detailed in Table S7.

A recent candidate gene study investigating the LCP6 variant (rs61764370) in the 3’-UTR of KRAS showed very strong association with endometriosis among 132 women (MAFcase = 0.311, MAFcontrol = 0.076) [22], and proposed that let-7 microRNA play a functional role in the development of endometriosis. We investigated this association by Taqman assay in a set of 1123 cases and 832 controls from our European population and found the allele frequencies to be identical in the two populations (MAFcase = 0.095, MAFcontrol = 0.093). Our result show that the LCS6 variant of KRAS does not provide any value as an indicator for endometriosis risk in a European population in agreement with results reported by Luong et al. [23].


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Clinical Stratification and Diagnostic Delay

Clinical features commonly used to characterize and stratify endometriosis include infertility, pelvic pain, severity, age-at-menarche and familiality. To determine the diagnostic delay in our patient-population we identified a group of women (n = 874) that reported both age at onset-of-symptoms (mean-age-onset = 19.04 years) and age at diagnosis (mean-age-diagnosis = 27.49), and observed an average diagnostic delay of 8.44 years, similar to previous studies. We then went on to examine if our samples showed any clinical correlations using logistic regression. The analysis revealed strong correlations between severity and infertility (P<0.001, OR = 2.19), and between severity and diagnostic delay (P<0.001, OR = 1.04) as shown
in Table 2. To identify loci associated with the progression of endometriosis, we compared patients with mild endometriosis to patients with moderate or severe endometriosis in a two-stage GWAS. Stage one included 657 patients with mild endometriosis vs. 529 patients with moderate-or-severe endometriosis and a stage two replication set of 319 mild vs. 519 moderate-or-severe patients. A meta-analysis using the CMH test in this limited sample set, found no loci that pass the genome-wide significant threshold which suggest the SNPs identified in the primary study contribute to the general endometriosis risk rather than endometriosis progression. A separate analysis of the top five SNPs using logistic regression also showed no correlation with severity as shown in Table S5a, and Table S5b shows no noticeable increase in effect size when comparing moderate and severe disease against all controls.

Risk Analysis of Endometriosis

After removing markers with \( r^2 > 0.8 \) among the top 5 associated regions (incl. the IL33 locus), we conducted multivariate logistic regression using the combined set of 2,019 cases and 14,471 controls. All of the 5 SNPs rs101917151, rs6757804, rs6907340, rs10975519 and rs10508881 analyzed remained significant with OR of 1.3, 1.2, 1.18, 1.17 and 1.17. Each marker appear to be an independent risk factors for endometriosis. Comparison of the OR between the discovery and replication datasets, shown in Table 1, does not suggest any significant inflation of effect size in the discovery dataset (winner’s curse), but this conclusion remain tentative due to the difference in size when comparing moderate and severe disease against all controls.

Conclusion

A two-stage GWAS and a replication study involving 2,019 cases and 14,471 controls was performed which identified four novel loci strongly associated with endometriosis and confirmed the involvement of a region around WNT4 which previously have been suggested as being associated to endometriosis. Nine other regions identified in the study also hold promise as candidate loci for endometriosis. Utmost care was taken in the clinical classification of patients and only surgically-confirmed cases with >95% European ancestry were considered in this large GWAS of endometriosis. The study is well powered (>90%) to identify a marker at or above 10% minor allele frequency (MAF) with odds-ratio (OR) >1.20, but we estimate the top 5 loci only explain about 1.5% of the phenotypic variance of endometriosis. Since the few risk loci we detected all have odds ratios <1.30 it must be assumed that any new endometriosis loci that contribute to the “missing heritability” must be rare, recent, or show minimal effect. GWAS, by design, detects only very old founder effects. When a phenotype includes infertility, like endometriosis, a high mutation rate would be required to replenish the disease-causing alleles lost from the gene-pool due to infertility. One suitable avenue to investigate under that scenario is to use whole genome sequencing of high-risk families rather than SNP-based GWAS.

Little is presently known about the pathophysiology of endometriosis, but we hope that a more detailed investigation of the loci presented in this paper will help elucidate the pathogenesis of endometriosis and clarify its genetic underpinnings.

Materials and Methods

Ethics Statement

All subjects and controls provided written informed consent in accordance with study protocols approved by Quorum Review IRB (Seattle, WA 98101).

Participant Recruitment

Patients included in the present study were invited to participate via an outreach program at www.endoendo.com, where our research initiative is described in more detail. Briefly, the “End to Endo” website provides general information regarding endometriosis and our research project, and invites women diagnosed with endometriosis to participate in our study.

Medical Review

The inclusion criteria in the endometriosis case population in the present study is surgically confirmed diagnosis of endometriosis with laparoscopy being the preferred method. Trained OB/GYN clinicians performed the medical record review and clinical assessment of each individual patient. Patients were considered to be affected if they had biopsy-proven lesions or if operative reports revealed unambiguous gross lesions. Patients were further categorized by severity, clinical history of pelvic pain, infertility, dyspareunia or dysmenorrhea and family history. Patients were grouped into one of three classes of severity: mild, moderate or severe, following the general guidelines set forth by ASRM [25].

Exclusion of endometriosis also requires surgical intervention and we made no attempt to exclude endometriosis in the population controls. Thus, in this analysis we are comparing cases with 100% prevalence of endometriosis to controls with the population prevalence of endometriosis (5–10%), which leads to a systematical underestimation of the true odds ratios and a decrease in statistical power to detect associations.

DNA Extraction

Saliva samples were collected using the Oragene 300 saliva collection kit (DNA Genotek; Ottawa, Ontario, Canada) and

| Table 2. Endometriosis severity correlate with infertility and diagnostic delay. |
|-----------------|-----------------|-----------------|------------|---|---|---|---|
| Clinical Feature | Moderate or Severe endometriosis (n=842) | Mild endometriosis (n=1177) | Category | OR | Beta | SE | P       |
| Infertility (1182) | 525 | 657 | Yes or No | 2.19 | 0.78 | 0.12 | 8.52E-11 |
| Family History (1881) | 790 | 1091 | Yes or No | 0.89 | -0.11 | 0.09 | 0.23 |
| Age at Menarche (921) | 405 | 516 | <=12 yrs or >12 yrs | 1.20 | 0.18 | 0.13 | 0.18 |
| Diagnostic Delay (874) | 383 | 491 | 0 to 35 | 1.04 | 0.04 | 0.01 | 2.18E-05 |

Clinical features were correlated to severity. Only patients that could be unambiguously categorized were included in the analysis with total counts provided in parenthesis next to the clinical feature. P-values (P) are calculated using Wald test. Beta is the regression coefficients and SE the standard error from logistic regression.

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DNA was extracted using an automated extraction instrument, AutoPure LS (Qiagen; Valencia, CA), and manufacturer’s reagents and protocols. DNA quality was evaluated by calculation absorbance ratio OD$_{260}$/OD$_{280}$, and DNA quantification was measured using PicoGreen® (Life Technologies; Grand Island, NY).

**Microarray Genotyping**

The discovery set of 1514 cases and 12660 controls and replication set of 505 cases and 1011 controls were genotyped using the Illumina Human OmniExpress Chip (Illumina; San Diego, CA) according to protocols provided by the manufacturer. Figure S4 show the genotype clusters for the top eight SNPs in our study. All SNPs reported in the present study passed visual inspection for cluster quality. It is our experience that technical replication does not affect genotype calls of SNPs with high quality clusters and due to cost we could not justify independent technical replication.

**Taqman Genotyping**

A Taqman® 7900 instrument (Life Technologies; Grand Island, NY) and manufacturer’s protocols were used to genotype rs61764370. Genotypes were determined using Taqman genotyping software SDS (v2.3) and the genotype cluster was visually inspected. Genotyping QC for rs61764370 passed standard criteria of call rate >95% and no deviation from HWE ($p<0.001$) was observed.

**Sample Quality Control**

Samples were excluded from the analysis if they missed any of the following quality thresholds:

- a) Evidence of familial relationship closer that 3rd-degree ($\pi>0.2$) using genome-wide Identity-By-State (IBS) estimation implemented in PLINK
- b) Samples with missing genotypes >0.02
- c) Samples with non-European admixture >0.05 as determined by ADMIXTURE

**SNP Quality Control**

SNPs were excluded from the analysis if they missed any of the following quality thresholds:

- a) SNPs with Illumina GenTrain Score <0.65
- b) SNPs from copy number variant regions or regions with adjacent SNPs
- c) SNPs failing Hardy-Weinberg Equilibrium (HWE) $P\leq10^{-3}$
- d) SNPs with minor allele frequency (MAF) $\leq0.01$ in the control population
- e) SNP call rate $\leq98\%$

**Admixture**

ADMIXTURE (ver. 1.22) was used to estimate the individual ancestry proportion [16]. The software estimates the relative admixture proportions of a given number of a priori defined ancestral groups contributing to the genome of each individual. We used the POPRES dataset [26] as a reference group to create a supervised set of 9 ancestral clusters. Seven of them belong to the European subgroups along with African and Asian groups. Since POPRES dataset utilized Affymetrix 5.0 chip, we used 105,079 autosomal SNPs that overlapped with the Illumina OmniExpress dataset. Among the 105,079 SNPs we selected a subset of 33,067 SNPs that showed greater genetic variation (absolute difference in frequency) among the 9 reference groups. The pair-wise autosomal genetic distance determined by Fixation Index ($F_{ST}$) using 33,067 SNPs was calculated for the 9 reference groups and show in Table S6 [27]. Subsequently, a conditional test was used to estimate the admixture proportions in the unknown samples as described by Alexander et al. (2009).

**Principal Component Analysis (PCA)**

PCA was applied to account for population stratification among the European subgroups. We selected the previously identified 33,067 SNPs to infer the axes of variation using EIGENSTRAT [28]. Only the top 10 eigenvectors were analyzed. Most of the variance among the European populations was observed in the first and second eigenvector. The first eigenvector accounts for the east-west European geographical variation while the second accounts for the north-south component. Only the top 10 eigenvectors showed population differences using Anova statistics ($p<0.01$). We then calculated the PCA adjusted Armitage trend P-values using the top 10 eigenvectors as covariates.

**Power Analysis**

Power calculations were performed using QUANTO (ver. 1.2), using a log-additive model. The analysis included 2019 cases and 14471 controls with the following assumptions: Type I error = 0.05, a minor allele frequency $\geq0.10$ and the odds-ratio $\geq1.2$.

**Association Analysis**

After the quality of all data was confirmed for accuracy, genetic association was determined using the whole-genome association analysis toolset, PLINK (ver. 1.07) [29].

Differences in allele frequencies between endometriosis patients and population controls were tested for each SNP by a 1-degree-of-freedom Cochran-Armitrage Trend test.

The allelic odds ratios were calculated with a confidence interval of 95%. SNPs that passed the quality control parameters were used to calculate the genomic inflation factor ($\lambda$) as well as to generate Quantile-Quantile (QQ) plots (Figure S2), which were generated by ranking a set of –log10 P-values and plotting them against their expected values. PCA adjusted Cochran-Armitrage trend test P-values were also determined. The combined/meta-analysis of discovery and replication dataset was performed using Cochran-Mantel-Hanszel method. Breslow Day test was used to determine between-cluster heterogeneity in the odds ratio for the disease/SNP association. Multivariate Logistic regression was used to test for independence of SNP effects. Univariate Logistic regression was used to test for correlation of clinical factors to the severity of the disease.

Control samples include both male and female samples in approximately equal proportions. The allele frequencies for the 8 strongly associated SNPs and the 15 SNPs with suggested associations did not show any significant gender bias.

Haplotype-based association tests were calculated by 1-degree of freedom $\chi^2$-test, along with their respective odds ratios using PLINK.

The variance explained by logistic regression model is calculated using the Cox Snell and Nagelkerke pseudo $R^2$ method which is similar to the $R^2$ concept of linear regression [30].

**Imputation Analysis**

IMPUTE2 (ver. 2.2.2) was used for imputing SNPs against the 1000-Genome (version 3 of the Phase 1 integrated data). Samples
were pre-phased with IMPUTE2 using actual genotypes and then imputed for SNPs included in the 1000-Genome reference panel to form imputed haplotypes. Imputation was carried out within +/- 250 kb of the main marker of interest. Only SNPs that pass the confidence score of > 0.9 from imputation, call rate of 0.95 and with MAF>0.01 are reported. The imputation was performed on the total dataset of 2,019 cases and 14,471 control subjects.

Software Used

PLINK (version 1.07; http://pngu.mgh.harvard.edu/~purcell/plink/index.shtml).
QUANTO (version 1.2; http://hydra.usc.edu/gxe).
R (version 2.15.0; http://www.r-project.org/).
Impute2 (version 2.2.2; http://mathgen.stats.ox.ac.uk/impute/impute_v2.html).
LocusZoom (version 1.1; http://csg.sph.umich.edu/locuszoom/) was used for regional association plots.

EIGENSTRAT (version 3.0; http://genepath.med.harvard.edu/~reich/Software.htm).

Supporting Information

Figure S1 PCA classification of Case and Control samples. A reference set of samples previously identified as European are shown in Panel A (reference). Samples selected for being 95% European are projected onto the European map and shown in Panel B. The Figure show that our Case and Control populations are geographically identical. A preponderance of the European are shown in Panel A (reference). Samples selected for being 95% European are projected onto the European map and shown in Panel B. The Figure show that our Case and Control populations are geographically identical. A preponderance of the participants have ancestral roots in the north-western part of Europe with a Southern trend towards Italy.

Figure S2 Quantile-quantile plots for the Discovery set of 1,514 endometriosis cases and 12,660 population controls before and after PCA-based adjustment. The unadjusted QQ plot in Panel A is showing a \( \lambda = 1.18 \). The adjusted QQ plot in Panel B show \( \lambda = 1.05 \). The association analysis include 580,699 SNPs and included only samples that passed our Ethnicity, SNP and Sample quality filters.

Figure S3 Regional association plots for the five top regions (Panel A–E).

Figure S4 Genotype clusters for the 8 most strongly associated SNPs.

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