BIG BIASES AMONGST PRODUCTS OF TWO PRIMES

DAVID DUMMIT, ANDREW GRANVILLE, AND HERSHY KISILEVSKY

Abstract. We show that substantially more than a quarter of the odd integers of the form $pq$ up to $x$, with $p, q$ both prime, satisfy $p \equiv q \equiv 3 \pmod{4}$.

1. Introduction

There are roughly equal quantities of odd integers $n$ that are the product of two primes, $p$ and $q$, in the two arithmetic progressions $1 \pmod{4}$ and $3 \pmod{4}$. Indeed the counts differ by no more than $x^{1/2+o(1)}$ (assuming the Riemann Hypothesis for $L(1,(-4/\cdot))$; see [1] for a detailed analysis). One might guess that these integers are further evenly split amongst those with $(p \pmod{4}, q \pmod{4}) = (1,1)$, $(1,-1)$, $(-1,1)$ or $(-1,-1)$, but recent calculations reveal a substantial bias towards those $pq \le x$ with $p \equiv q \equiv 3 \pmod{4}$. Indeed for the ratio

$$r(x) := \frac{\# \{pq \le x : p \equiv q \equiv 3 \pmod{4}\}}{\frac{1}{4}\# \{pq \le x\}}$$

we found that

$$r(1000) \approx 1.347, \quad r(10^4) \approx 1.258, \quad r(10^5) \approx 1.212, \quad r(10^6) \approx 1.183, \quad r(10^7) \approx 1.162,$$

a pronounced bias that seems to be converging to 1 surprisingly slowly. We will show that this is no accident and that there is similarly slow convergence for many such questions:

Theorem 1.1. Let $\chi$ be a quadratic character of conductor $d$. For $\eta = -1$ or $1$ we have

$$\frac{\# \{pq \le x : \chi(p) = \chi(q) = \eta\}}{\frac{1}{4}\# \{pq \le x : (pq, d) = 1\}} = 1 + \eta \left( \frac{\mathcal{L}_\chi + o(1)}{\log \log x} \right) \quad \text{where} \quad \mathcal{L}_\chi := \sum_p \frac{\chi(p)}{p}.$$

If $\chi = (-4/\cdot)$ then $\mathcal{L}_\chi = -3.34\ldots$ so the theorem implies that $r(x) \ge 1 + \frac{(1+o(1))}{3(\log \log x - 1)}$. If we let $s(x) = 1 + \frac{1}{3(\log \log x - 1)}$ then we have

$$s(1000) \approx 1.357, \quad s(10^4) \approx 1.273, \quad s(10^5) \approx 1.230, \quad s(10^6) \approx 1.205, \quad s(10^7) \approx 1.187,$$

a pretty good fit with the data above. The prime numbers have only been computed up to something like $10^{24}$ so it is barely feasible that one could collect data on this problem up to $10^{50}$ in the foreseeable future. Therefore we would expect this bias to be at least 7% on any data that will be collected this century (as $s(10^{50}) \approx 1.07$).
Proof. For a given quadratic Dirichlet character \( \chi \) we will count the number of integers \( pq \leq x \) with \( \chi(p) = \chi(q) = 1 \) (and the analogous argument works for \(-1\)). One can write any such integer \( pq \leq x \) with \( p \leq q \leq x/p \), so that \( p \leq \sqrt{x} \). Hence we wish to determine

\[
\sum_{p \leq \sqrt{x}} \sum_{\substack{q \leq x/p \\chi(p) = 1 \\chi(q) = 1}} 1.
\]

The prime number theorem for arithmetic progressions reveals that \( \sum_{q \leq Q, \chi(q) = 1} 1 = \frac{Q}{2 \log Q} + O\left(\frac{Q}{(\log Q)^2}\right) \), so the above sum equals

\[
\sum_{p \leq \sqrt{x}} \left( \frac{\chi_0(p) + \chi(p)}{2} \right) \frac{x}{2p \log(x/p)} + O\left(\frac{x}{p \log x} + \frac{p}{\log p}\right).
\]

where the implicit constant in the \( O(.) \) depends only on the conductor of \( \chi \). This equals

\[
\frac{1}{4} \sum_{p \leq \sqrt{x}} \frac{x}{p \log(x/p)} + \frac{x}{4} \sum_{p \leq \sqrt{x}} \frac{\chi(p)}{p \log(x/p)} + O\left(\frac{x}{(\log x)^2 \log \log x}\right).
\]

The difference between the second sum, and the same sum with \( \log(x/p) \) replaced by \( \log x \), is

\[
\frac{x}{4 \log x} \sum_{p > \sqrt{x}} \frac{\chi(p)}{p \log(x/p)} \ll \frac{x}{(\log x)^2},
\]

using the prime number theorem for arithmetic progression and partial summation. Moreover

\[
\frac{x}{4 \log x} \sum_{p > \sqrt{x}} \frac{\chi(p)}{p} \ll \frac{x}{(\log x)^2}.
\]

Collecting together what we have proved so far yields that \#\{pq \leq x : \chi(p) = \chi(q) = 1\}

\[
= \frac{1}{4} \left\{ \#\{pq \leq x : (p, d) = 1\} + \frac{x}{\log x} \sum_p \frac{\chi(p)}{p} + O\left(\frac{x}{(\log x)^2 \log \log x}\right) \right\}
\]

The first term is well-known to equal \( \frac{x}{\log x} (\log \log x + O(1)) \), and so we deduce that

\[
\frac{4\#\{pq \leq x : \chi(p) = \chi(q) = 1\}}{\#\{pq \leq x : (p, d) = 1\}} = 1 + \frac{1}{\log \log x} \left( \sum_p \frac{\chi(p)}{p} + o(1) \right).
\]

as claimed. \( \square \)

We note that

\[
\sum_p \frac{\chi(p)}{p} = \sum_{m \geq 1} \frac{\mu(m)}{m} \log L(m, \chi^m) = \log L(1, \chi) + E(\chi),
\]

where

\[
\sum_p \left( \log \left( 1 - \frac{1}{p} \right) + \frac{1}{p} \right) = -0.315718 \ldots \leq E(\chi) \leq \sum_p \left( \log \left( 1 + \frac{1}{p} \right) - \frac{1}{p} \right) = -0.18198 \ldots
\]
2. Further remarks

- One deduces from our theorem that \( r(x) > 1 \) for all \( x \) sufficiently large and we conjecture that this is true for all \( x \geq 9 \).
- We also conjecture that \( \mathcal{L}_\chi \) is always non-zero so that there is always such a bias.
- One can calculate the bias in other such questions. For example, we get roughly triple the bias for the proportion of \( pq \leq x \) for which \( \left( \frac{5}{p} \right) = \left( \frac{5}{q} \right) = -1 \) out of all \( pq \leq x \) with \( p, q \neq 5 \) (since \( \mathcal{L}_{\chi/5} \approx -1.008 \)). The data
  \[
  r_5(1000) \approx 1.881, \quad r_5(10^4) \approx 1.626, \quad r_5(10^5) \approx 1.523, \quad r_5(10^6) \approx 1.457, \quad r_5(10^7) \approx 1.416,
  \]
  confirms this very substantial bias. It would be interesting to find more extreme examples.
- How large can the bias get if \( d \leq x \)? It is known \([2]\) that \( L(1, \chi) \) can be as large as \( c \log \log d \), and so \( \mathcal{L}_\chi \) can be as large as \( \log \log \log d + O(1) \). We conjecture that there exists \( d \leq x \) for which the bias in our Theorem is as large as
  \[
  1 + \frac{\log \log \log x + O(1)}{\log \log x}.
  \]
  Note that this requires proving a uniform version of the Theorem. Our proof assumes that \( x \) is allowed to be very large compared to \( d \), so does not immediately apply to the problem that we have just stated.
- The same bias can be seen (for much the same reason) in looking at
  \[
  \sum_{\substack{p \leq x \\text{\(p \equiv 3\) (mod 4)}}} \frac{1}{p} / \sum_{\substack{p \leq x \\text{\(p \equiv 1\) (mod 4)}}} \frac{1}{p} \approx 1 + \frac{2}{3 \log \log x}.
  \]
  Indeed, by the analogous proof, we have in general
  \[
  \sum_{\substack{p \leq x \\chi(p) = 1}} \frac{1}{p} / \sum_{\substack{p \leq x \\chi(p) = -1}} \frac{1}{p} = 1 + 2 \frac{(\mathcal{L}_\chi + o(1))}{\log \log x}.
  \]
  We therefore see a bias in the distribution of primes in arithmetic progressions, where each prime \( p \) is weighted by \( 1/p \), as a consequence of the sign of \( \mathcal{L}_\chi \). This effect is much more pronounced than in the traditional prime race problem where the same comparison is made, though with each prime weighted by 1. The bias here is determined by the distribution of values of \( \chi(p) \), whereas the prime race bias is determined by the values of \( \chi(p^2) = 1 \), so they appear to be independent phenomena. However one might guess that both biases are sensitive to low lying zeros of \( L(s, \chi) \). This probably deserves further investigation, to determine whether there are any correlations between the two biases.
- One can show the following for \( k \) prime factors, by similar methods:
  \[
  \frac{\# \{ p_1 \ldots p_k \leq x : \text{each } \chi(p_j) = \eta \}}{\frac{1}{x} \# \{ p_1 \ldots p_k \leq x : \text{each } (p_j, d) = 1 \}} = 1 + \eta \frac{((k - 1)\mathcal{L}_\chi + o(1))}{\log \log x}.
  \]
  It would be interesting to understand this when \( k \) gets large, particularly when \( k \sim \log \log x \), the typical number of prime factors of an integer \( \leq x \). It seems likely that the factor on the
right-side should grow like
\[ c_{\chi,k} \left( 1 + \eta \frac{\mathcal{L}_\chi + o(1)}{\log \log x} \right)^{k-1}, \]
but we do not know what \( c_{\chi,k} \) would look like.

- More generally, if \( \chi_1, \ldots, \chi_k \) are quadratic characters (with \( \chi_j \) of conductor \( d_j \), and each \( \eta_j = -1 \) or \( 1 \) then
\[
\frac{\# \{ p_1 \ldots p_k \leq x : \chi_j(p_j) = \eta_j \text{ for each } j \}}{\frac{1}{2^k} \# \{ p_1 \ldots p_k \leq x : \text{each } (p_j, d_j) = 1 \} } = 1 + \frac{((k-1)c(\vec{\chi}, \vec{\eta}) + o(1))}{\log \log x}.
\]
where
\[ c(\vec{\chi}, \vec{\eta}) := \frac{1}{k} \sum_{j=1}^k \eta_j L_{\chi_j}. \]
In particular this type of bias does not appear when \( k = 2 \), \( \chi_1 = \chi_2 \) and \( \eta_1 + \eta_2 = 0 \). Can one prove that \( c(\vec{\chi}, \vec{\eta}) \) can only be 0 for such trivial reasons? That is, is \( c(\vec{\chi}, \vec{\eta}) = 0 \) if and only if \( \sum_j : x_j = \chi \eta_j = 0 \) for every character \( \chi \in \vec{\chi} \)?

- Given arithmetic progressions \( a \pmod{m} \) and \( b \pmod{n} \), one can surely prove that there exists \( \beta = \beta(a \pmod{m}, b \pmod{n}) \) such that
\[
\frac{\# \{ pq \leq x : p \equiv a \pmod{m}, q \equiv b \pmod{n} \}}{\phi(m)\phi(n)} \frac{\# \{ pq \leq x : (p, m) = (q, n) = 1 \} }{\log \log x} = 1 + \frac{\beta + o(1)}{\log \log x}.
\]
It would be interesting to classify when \( \beta(a \pmod{m}, b \pmod{n}) \) is non-zero, and to determine situations in which it is large. Or more generally for what subsets \( A \subseteq (\mathbb{Z}/m\mathbb{Z})^\ast \) and \( B \subseteq (\mathbb{Z}/n\mathbb{Z})^\ast \) is there no such bias? We would guess that this would only be the case if either

(i) \( A \) and \( B \) both contain all congruence classes (that is, every prime not dividing \( mn \) can be represented by both \( A \) and \( B \)); or

(ii) \( A \cup B \) is a partition of the integers coprime to \( mn \) (that is, every prime not dividing \( mn \) is represented by \( A \), or represented by \( B \), but not both).

References

[1] Kevin Ford and Jason Sneed, *Chebyshev’s bias for products of two primes*, Experiment. Math. 19 (2010), 385-398.
[2] Andrew Granville and K. Soundararajan, *The distribution of values of \( L(1, \chi_d) \)*, Geom. Funct. Anal. 13 (2003), 992-1028.

Department of Mathematics and Statistics, University of Vermont, Burlington, VT 05401, USA

E-mail address: dummit@math.uvm.edu

Centre de recherches mathématiques, Université de Montréal, CP 6128 succ. Centre-Ville, Montréal, QC H3C 3J7, Canada

E-mail address: andrew@dms.umontreal.ca

Department of Mathematics and Statistics, Sir George Williams Campus, Concordia University, Montreal, QC H3G 1M8, Canada

E-mail address: kisilev@mathstat.concordia.ca