Examining within-person relationships between state assessments of affect and eudaimonic well-being using multi-level structural equation modeling

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ABSTRACT
Prior research has highlighted the possibility that current affect may be interchangeable with state assessments of well-being at the within-person level (E. Jayawickreme et al., 2017a). While this finding points to the potential disutility of measuring distinct subjective well-being (SWB) states, it is unclear whether this extends to other dimensions of well-being. Here we present the results of a systematic replication and extension of E. Jayawickreme et al. (2017a), which employed a multi-level structural equation model examination of the within-person relationship between state assessments of affect and life satisfaction. This study found that measures of state affect and satisfaction were redundant in the sense that within-person changes in state affect perfectly explained within-person changes in state satisfaction. The goal of the present investigation was to ascertain whether state assessments of key dimensions of well-being (EWB) provide distinct information about individuals’ well-being over and above their current affect.

Assessing well-being
Psychologists have assessed well-being in many forms over the last half-century (E. Jayawickreme et al., 2012). One consistent distinction in this literature is that between the affective and cognitive components of SWB (Diener, 1984) on the one hand, and EWB on the other. While a variety of definitions of EWB exist in the literature (Jayawickreme & Dahill-Brown, 2016), a unifying theme is that EWB distinctively reflects positive functioning that emanates from the possession of key dimensions of human thriving (e.g., the psychological well-being [PWB] perspective; Ryff, 1989). These dimensions include the ability to choose and create environments that fit one’s personality, having positive interpersonal relationships, holding positive self-evaluations of one’s past and present self (i.e., self-acceptance), developing and growing as a person, and possessing self-determination, independence, self-regulation of behavior, meaning and purpose in life, and authenticity (Jayawickreme & Dahill-Brown, 2016; Ryff, 1989; A. S. Waterman, 1993, 2008).

These forms of well-being can be assessed with both trait and state (experiential) measures. Global or trait well-being measures are posited to assess individuals’ beliefs about the overall well-being of their lives, whereas state well-being measures aim to assess short-term or momentary reports of well-being in the context of lived experiences (Hudson et al., 2017). Notably, previous research has indicated that reports of state well-
being are influenced by appraisals of global life circumstances (Luhmann et al., 2012).

**Examining the role of affect in well-being judgements**

A major contemporary issue relevant to the valid measurement of state well-being is whether individuals, when reporting on state well-being, rely on information unique to well-being (involving global evaluations) or instead rely merely on their current affect (see E. Jayawickreme et al., 2017a, 2017b for a more in-depth discussion). In other words, if someone reports their current level of well-being, to what extent would they use their current level of affect to guide their answer? We define this ‘effect of affect’ on well-being as the proportion of total variance in well-being explained by changes in affect over time. For life satisfaction, the effect using a trait measure assessed weekly was relatively inconsequential as there was relatively little within-person variance in trait life satisfaction (12; 9% of the total variance) from week to week over a five week period (E. Jayawickreme et al., 2017b). In contrast, the ‘effect of affect’ was more consequential for daily state measures of life satisfaction due to greater within-person variance over a two-to-three week period (87; 49% of the total variance); current affect accounted for 100% of the within-person variance in state satisfaction assessments (E. Jayawickreme et al., 2017a).

These findings suggest that there is an ‘effect of affect’ that is specific to daily or state assessments of life satisfaction and raise the possibility that current affect may have a substantial impact on state assessments of other dimensions of well-being. This is because whereas past research has argued for a clear distinction between SWB and PWB on conceptual grounds (e.g., Ryff, 1989), existing research has found support for a link between affect and state meaning (e.g., King et al., 2006) as well as for between mood and state authenticity (Lenton et al., 2013). More generally, some researchers have questioned the extent to which SWB and EWB are in fact empirically distinguishable (e.g., Disabato et al., 2016; Kashdan et al., 2008).

These considerations lead to the following question: does current affect explain the within-person variability in state assessments of EWB (such that measuring state EWB would be interchangeable with measuring current affect), or does state EWB provide unique information on state well-being over and above current affect? We answer this question here through a systematic replication and extension of E. Jayawickreme et al. (2017a) where we examined the impact of affect on state assessments of four distinct dimensions of EWB: meaning, core self-evaluation (the fundamental appraisals individuals make about their self-worth and capabilities), authenticity, and gratitude. In doing so, we examine whether current affect impacts state assessments of EWB to the same extent as it impacts state satisfaction.

**Using multi-level structural equation models (ML-SEM) to assess within-person relationships**

We note that we use multi-level structural equation models (ML-SEM) that can distinguish between ‘state’ (i.e., deviations from an individual’s average across time) and ‘trait’ (i.e., stable between-individual differences) variance for these analyses. Baumert et al. (2017) have highlighted the importance of integrating person-ality process, structure, and development, and answering such questions necessitate the examination of dynamics of trait change, intrapersonal variability of states and traits, and the interrelations between states and traits (Kandler, 2017) at both the fine-grained and long-term level (Finnigan & Vazire, 2017).

To assess personality at the within-person state level, it is important to use appropriate measures and statistical models (Brose et al., 2020). Prior research has mostly examined the affect-well-being relationships at the between-person level (e.g., the proportion of overlapping variance between affect and life satisfaction; Diener et al., 2012; Kuppens et al., 2008; Suh et al., 1998). One limitation of this work is that once a relationship among two variables is identified at the global (between-person) level, it is often inferred that the relationship holds true at the momentary (within-person) level (i.e., the ecological fallacy). As discussed by E. Jayawickreme et al. (2017a, 2017b), within-person approaches may reveal different answers than between-person approaches, since the causes for why variables may vary across people may be different from why they vary within a person across situations. Many key questions related to the impact of states on traits are fundamentally within-person questions, and the direction and magnitude of any identified impact will vary depending on the level of analysis (Molenaar & Campbell, 2009).

Using multi-level structural equation modeling (ML-SEM) to address questions regarding (1) partitioning total variance into between-person ‘trait’ (i.e., stable between-individual differences) and within-person ‘state’ (i.e., deviations from an individual’s average across time) components, and (2) estimating within-person relationships between state variables holds unique
advantages. The multilevel aspect of the ML-SEM allows the simultaneous analysis and comparison of the within- and between-person relationships among variables, whereas the structural equation modeling aspect enables the separation of error variance from reliable latent constructs. Without this technique, error variance would be confounded with within-person variance. Moreover, the critical question of whether reports of well-being largely assess current affect is fundamentally a within-person question. In other words, do changes in affect impact changes in well-being for a particular person? For example, if an individual’s current positive affect increases, will her EWB rating also increase?

The ML-SEM approach employed here is also different from more classical state-trait models (e.g., Eid & Diener, 2004). State-trait models use a residual-based approach that addresses the ecological fallacy discussed above by removing overall between-person variance from the occasion-specific effects. In other words, the occasion-specific variables are residuals (or deviation scores) that have been adjusted/control for the trait (i.e., an individual’s average over time), and these models examine effects across individuals within occasions (e.g., does a person with a relatively large occasion-specific residual of well-being also have a relatively large occasion-specific residual of affect relative to other individuals?). However, this approach is conceptually complicated and relatively difficult to interpret/comprehend. Our approach is conceptually more straightforward/direct (where the impact of affect is basically the within-person relationship between affect and EWB).

The Present Study
In the present study, we evaluated state assessments of four dimensions of EWB: meaning, core self-evaluation, authenticity, and positive relationships.

Meaning
Meaning has been seen in the psychological literature as distinct from happiness (Baumeister, 1991; Hicks & King, 2007; Ryff & Keyes, 1995; c.f. Disabato et al., 2016). Typically, whereas measures of subjective well-being are seen as contributing to hedonia, meaning is seen as a central dimension of eudaimonia (i.e., well-being derived not from pursuing momentary desires but rather experiences that promote growth and wellness) (Dahill-Brown & Jayawickreme, 2016; Ryan & Deci, 2001). Of note, the present work replicates and extends previous examinations of relations between affect and daily meaning (e.g., King et al., 2006) while employing a novel statistical approach. Specifically, our models produce separate estimates of both the between- and the within-person impact of affect on meaning, whereas prior work (e.g., Study 2 in King et al., 2006) did not distinguish between them, which produces a confounded amalgam of the two effects. Furthermore, our models use latent variables to adjust for measurement error.

Core self-evaluation
Core self-evaluations refer to the fundamental appraisals individuals make about their self-worth and capabilities (Bono & Judge, 2003). The constructs of self-esteem, neuroticism, locus of control, and generalized self-efficacy have typically made up the key dimensions of this construct (Judge et al., 2002). Notably, they capture key features of meaningful goal-directedness characteristic of EWB theories (specifically, holding positive self-evaluations of oneself, and possessing qualities such as self-determination, independence, self-regulation of behavior; Ryff, 1989).

Authenticity
Authenticity, or the sense that one generally acts in a manner congruent with one’s true self, has been seen by some theories as a central dimension of EWB (A. S. Waterman, 1993, 2013; Sheldon & Kasser, 2001). State authenticity has been defined as the sense that one is currently in alignment with one’s true or genuine self (Sedikides et al., 2017). It is frequently reported in daily life (Lenton et al., 2013) and is associated in increased subjective vitality (Thomaes et al., 2017) and meaning (Schlegel et al., 2009). Moreover, experimental work has suggested that people use current mood information in evaluating their state authenticity (Lenton et al., 2013).

Positive Relationships
Ryff (1989) identified positive relationships as a key dimension of EWB. Ryan and Deci (2001) have similarly identified relatedness – the importance of feeling a close connection to and being cared for by others – as one of three psychological needs that contribute to a flourishing life. The importance of positive relationships has been shown to be valued across cultures (Sheldon et al., 2001) and further relates to measures of life satisfaction (Reis et al., 2000; Sheldon et al., 1996).

Materials and Methods
Participants
Two hundred and seven participants were recruited from Amazon.com’s Mechanical Turk online recruitment
platform. Due to a coding error, demographic information for three of these participants were not recorded. The mean age of the 204 participants for which we obtained demographic information was 33.72 (SD = 10.78) with 109 identifying as men and 95 as women. In terms of self-reported racial background, 167 reported being White, 15 were Black, 17 were Asian, and 4 were from other racial backgrounds. Participants completed assessments of well-being at the end of each evening for a period of 14 consecutive days (see Table 1 for the list of item). They received a link to the survey (which was hosted on the Qualtrics survey platform) each evening via email.

**Measures**

**State meaning**

State meaning was assessed with four items adapted from the Meaning in Life questionnaire (Steiger et al., 2006). The estimated reliability was .98 at the between-person level and .94 at the within-person level.\(^1\)

**State affect**

Participants rated one item measuring positive affect (‘Today, I felt happy’), and one item measuring negative affect (‘Today, I felt sad’), following Kashdan and Steger (2007). These items were rated on a seven-point scale ranging from ‘Did not feel this way at all’ to ‘Felt this way very strongly.’

| Table 1. Items utilized. |
|--------------------------|
| Meaning | How meaningful did you feel your life was today? |
| | I did something today that made my life feel meaningful |
| | How much did you feel that your life had purpose today? |
| | I felt a sense of clarity about my life’s purpose today |
| Core self-evaluation | I felt confident about achieving my goals today |
| | I felt like my efforts were successful today |
| | I failed at something today, and I felt inadequate |
| | I completed my tasks successfully today |
| | I did not feel in control of my routine today |
| | I felt satisfied with myself today |
| | I had doubts about my competence today |
| | I felt that I was in control of my life today |
| | I successfully met my goals today |
| | I felt capable of coping with most of my problems today |
| Authenticity | I felt that I wore a number of social masks |
| | Today I felt that throughout the day I was in touch with my true self |
| | Today I felt that my choices expressed my true self |
| Positive Relationships | I had a meaningful conversation with a family member or friend today |
| | I felt distant from the people around me today |
| Current affect | Today, I felt sad |
| | Today, I felt happy |

**State core self-evaluation**

State core self-evaluation was assessed using 10 items adapted from the CORE Inventory (Judge et al., 2002). We did not include two emotion-related items that might artificially inflate the relationship of core self-evaluations with our emotion predictors: ‘I felt depressed today’ and ‘Today, I felt things were pretty bleak and hopeless.’ The estimated reliability was .99 at the between-person level and .98 at the within-person level.

**State authenticity**

State authenticity was assessed using three items adapted from Heppner et al. (2008). The estimated reliability was .93 at the between-person level and .89 at the within-person level.

**Positive Relationships**

State positive relationships were assessed with two items adapted from the Positive Relationships with Others subscale of the Psychological Well-Being Questionnaire (Ryff & Keyes, 1995). The estimated reliability of these items was .85 at the between-person level and .80 at the within-person level.

**Procedures and analytic approach**

We used multilevel structural question modeling (ML-SEM) with affect as a time-varying covariate in Mplus 7 (Muthén & Muthén, 2012). The multilevel aspect of our analyses allowed us to analyze and compare the within- and between-person relationships between affect and eudaimonic well-being simultaneously, whereas the structural equation modeling feature of our models enabled us to separate error variance from reliable latent constructs. Without this technique, error variance would be confounded with within-person variance. To scale the latent constructs, we used the effects-coding method of scaling (Little et al., 2006). This method allows the latent construct to remain in the same scale as the indicator variables. As a measure of effect-size, we used $r_{effect} = \sqrt{t^2/(t^2 + df)}$ (see Duckworth et al., 2010; Rosenthal & Rosnow, 1991, p. 441). We conducted a series of ML-SEM models. For each well-being outcome, Model 0 was an unconditional model with no predictors to estimate the amount of variance in well-being at each level (between- and within-person) and to serve as a baseline model against which to compare subsequent models including predictors. We added positive affect (Model 1) and negative affect (Model 2) individually and simultaneously (Model 3) as predictors of well-being at both the within-person...
and between-person levels. In Models 1–3, the within-person effects/slopes were allowed to vary across individuals (i.e., treated as random).

**Results**

**Meaning**

In Model 0 (no predictors), the variance estimates for daily meaning were 0.57 at the between-person level and 0.26 at the within-person level (see Table 2). The intraclass correlation (ICC) was .69 indicating that more than half of the variance was between-person (69%) vs. within-person (31%).

In Model 1 (positive affect), the between-person impact of positive affect was .82 (95% CI [.74, .90], \( r_{effect} = .81, p < .001 \)) and explained 66% of the between-person variance. At the within-person level, the impact of positive affect was .53 (95% CI [.48, .57], \( r_{effect} = .84, p < .001 \)) and explained 57% of the within-person variance in daily meaning.

In Model 2 (negative affect), the between-person impact of negative affect was −.48 (95% CI [−.58, −.37], \( r_{effect} = −.53, p < .001 \)) and explained 32% of the between-person variance. At the within-person level, the impact of negative affect was −.31 (95% CI [−.36, −.26], \( r_{effect} = −.67, p < .001 \)) and explained 35% of the within-person variance in daily meaning.

In Model 3 (both positive and negative affect), at the between-person level, the impact of positive affect was .80 (95% CI [.67, .92], \( r_{effect} = .65, p < .001 \)) and the impact of negative affect was −.03 (95% CI [−.13, .07], \( r_{effect} = −.04, p = .57 \)), together explaining 66% of the between-person variance in daily meaning. At the within-person level, the impact of positive affect was .46 (95% CI [.41, .51], \( r_{effect} = .82, p < .001 \)) and the impact of negative affect was −.11 (95% CI [−.14, −.08], \( r_{effect} = −.41, p < .001 \)) together explaining 60% of the within-person variance in daily meaning.

Using a state/daily measure of meaning, there was substantial within-person variance, about 31%. In addition, there was a substantial impact of affect for the state meaning measure (\( R^2 = .60 \)).

**Core Self-Evaluations**

In Model 0 (no predictors), the variance estimates for daily self-evaluations were 0.37 at the between-person level and 0.30 at the within-person level (see Table 3). The intraclass correlation (ICC) was .56 indicating that more than half of the variance was between-person (56%) vs. within-person (44%).

In Model 1 (positive affect), the between-person impact of positive affect was .61 (95% CI [.55, .67], \( r_{effect} = .83, p < .001 \)) and explained 86% of the between-person variance. At the within-person level, the impact of positive affect was .45 (95% CI [.41, .50], \( r_{effect} = .82, p < .001 \)) and explained 66% of the within-person variance in daily self-evaluation.

In Model 2 (negative affect), the between-person impact of negative affect was −.42 (95% CI [−.49, −.36], \( r_{effect} = −.65, p < .001 \)) and explained 69% of the between-person variance. At the within-person level, the impact of negative affect was −.35 (95% CI [−.39, −.31], \( r_{effect} = .74, p < .001 \)) and explained 61% of the within-person variance in daily self-evaluation.

In Model 3 (both positive and negative affect), at the between-person level, the impact of positive affect was .49 (95% CI [.43, .55], \( r_{effect} = .74, p < .001 \)) and the impact of negative affect was −.15 (95% CI [−.20, −.10], \( r_{effect} = −.40, p < .001 \)), together explaining 89% of the between-person variance in daily self-evaluation. At the within-person level, the impact of positive affect was .34 (95% CI [.30, .38], \( r_{effect} = .76, p < .001 \)) and the impact of

### Table 2. Multilevel structural equation model of affect predicting meaning.

|                      | Model 0 | Model 1 | Model 2 | Model 3 |
|----------------------|---------|---------|---------|---------|
| **Between-Individual** |         |         |         |         |
| Effects              |         |         |         |         |
| Positive Affect\(^a\) | .82 (.04)| .80 (.06)|         |         |
| Negative Affect\(^a\) | −.48 (.05)| −.03 (.05)|         |         |
| **Within-Individual** |         |         |         |         |
| Effects              |         |         |         |         |
| Positive Affect\(^b\) | .53 (.02)| .46 (.02)|         |         |
| Negative Affect\(^b\) | −.31 (.02)| −.11 (.02)|         |         |
| **Variance**         |         |         |         |         |
| Between-Individual   | 0.57    | 0.20    | 0.39    | 0.20    |
| Within-Individual     | 0.26    | 0.11    | 0.17    | 0.11    |
| \( R^2_{between} \) | 0.66    | 0.32    | 0.66    |         |
| \( R^2_{within} \)   | 0.57    | 0.35    | 0.60    |         |

Note. Standard errors are in parentheses. ICC = .69.

### Table 3. Multilevel structural equation model of affect predicting self-evaluations.

|                      | Model 0 | Model 1 | Model 2 | Model 3 |
|----------------------|---------|---------|---------|---------|
| **Between-Individual** |         |         |         |         |
| Effects              |         |         |         |         |
| Positive Affect\(^a\) | .61 (.03)| .49 (.03)|         |         |
| Negative Affect\(^a\) | −.42 (.03)| −.15 (.02)|         |         |
| **Within-Individual** |         |         |         |         |
| Effects              |         |         |         |         |
| Positive Affect\(^b\) | .45 (.02)| .34 (.02)|         |         |
| Negative Affect\(^b\) | −.35 (.02)| −.21 (.02)|         |         |
| **Variance**         |         |         |         |         |
| Between-Individual   | 0.37    | 0.05    | 0.12    | 0.04    |
| Within-Individual     | 0.30    | 0.10    | 0.12    | 0.08    |
| \( R^2_{between} \) | 0.86    | 0.69    | 0.89    |         |
| \( R^2_{within} \)   | 0.66    | 0.61    | 0.73    |         |

Note. Standard errors are in parentheses. ICC = .56.
negative affect was −.21 (95% CI [−.25, −.17], \( r_{\text{effect}} = .59, p < .001 \)) together explaining 73% of the within-person variance in daily self-evaluation.

Using a state/daily measure of core self-evaluation, there was substantial within-person variance, about 44%. In addition, there was a substantial impact of affect for the state self-evaluation measure (\( R^2 = .73 \)).

### Authenticity

In Model 0 (no predictors), the variance estimates for daily authenticity were 0.27 at the between-person level and 0.16 at the within-person level (see Table 4). The intraclass correlation (ICC) was 0.63 indicating that little more than half of the variance was between-person (63%) vs. within-person (37%).

In Model 1 (positive affect), the between-person impact of positive affect was .55 (95% CI [.45,.65], \( r_{\text{effect}} = .62, p < .001 \)) and explained 62% of the between-person variance. At the within-person level, the impact of positive affect was .41 (95% CI [.36,.47], \( r_{\text{effect}} = .72, p < .001 \)) and explained 56% of the within-person variance in daily authenticity.

In Model 2 (negative affect), the between-person impact of negative affect was −.33 (95% CI [−.43, −.24], \( r_{\text{effect}} = −.43, p < .001 \)) and explained 32% of the between-person variance. At the within-person level, the impact of negative affect was −.27 (95% CI [−.32, −.22], \( r_{\text{effect}} = −.60, p < .001 \)) and explained 40% of the within-person variance in daily authenticity.

In Model 3 (both positive and negative affect), at the between-person level, the impact of positive affect was .52 (95% CI [.41,.62], \( r_{\text{effect}} = .56, p < .001 \)) and the impact of negative affect was −.04 (95% CI [−.12, −.04], \( r_{\text{effect}} = −.07, p = .30 \)), together explaining 62% of the between-person variance in daily authenticity. At the within-person level, the impact of positive affect was .34 (95% CI [.29,.39], \( r_{\text{effect}} = .68, p < .001 \)) and the impact of negative affect was −.13 (95% CI [−.16, −.10], \( r_{\text{effect}} = −.47, p < .001 \)) together explaining 61% of the within-person variance in daily authenticity.

Using a state/daily measure of authenticity, there was substantial within-person variance, about 37%. In addition, there was a strong impact of affect for state authenticity (\( R^2 = .61 \)).

### Positive Relationships

In Model 0 (no predictors), the variance estimates for daily positive relationships were 0.31 at the between-person level and 0.27 at the within-person level (see Table 5). The intraclass correlation (ICC) was 0.53 indicating that a little more than half of the variance was between-person (53%) vs. within-person (47%).

In Model 1 (positive affect), the between-person impact of positive affect was .61 (95% CI [.51,.70], \( r_{\text{effect}} = .65, p < .001 \)) and explained 73% of the between-person variance. At the within-person level, the impact of positive affect was .44 (95% CI [.36,.52], \( r_{\text{effect}} = .59, p < .001 \)) and explained 58% of the within-person variance in daily positive relationships.

In Model 2 (negative affect), the between-person impact of negative affect was −.44 (95% CI [−.53, −.35], \( r_{\text{effect}} = −.58, p < .001 \)) and explained 45% of the between-person variance. At the within-person level, the impact of negative affect was −.33 (95% CI [−.35, −.31], \( r_{\text{effect}} = −.90, p < .001 \)) and explained 41% of the within-person variance in daily positive relationships.

In Model 3 (both positive and negative affect), at the between-person level, the impact of positive affect was .36 (95% CI [.23,.49], \( r_{\text{effect}} = .36, p < .001 \)) and the impact of negative affect was −.26 (95% CI [−.35, −.16], \( r_{\text{effect}} = −.35, p < .001 \)), together explaining 69% of the between-person variance in daily positive relationships. At the within-person level, the impact of positive affect

### Table 4. Multilevel structural equation model of affect predicting authenticity.

| Model 0 | Model 1 | Model 2 | Model 3 |
|---------|---------|---------|---------|
| **Between-Individual Effects** |          |         |         |
| Positive Affect*               | .55 (.05) | .52 (.05) |         |
| Negative Affect*               |          |         |         |
| **Within-Individual Effects**  |          |         |         |
| Positive Affect                | .41 (.03) | .34 (.03) |         |
| Negative Affect                |          |         |         |
| **Between-Individual Variance** | 0.27 | 0.10 | 0.09 |
| **Within-Individual Variance** | 0.16 | 0.07 | 0.10 |
| \( R^2_{\text{between}} \)    | 0.62 | 0.32 | 0.62 |
| \( R^2_{\text{within}} \)     | 0.56 | 0.40 | 0.61 |

Note. Standard errors are in parentheses. ICC = .63.

### Table 5. Multilevel structural equation model of affect predicting positive relationships.

| Model 0 | Model 1 | Model 2 | Model 3 |
|---------|---------|---------|---------|
| **Between-Individual Effects** |          |         |         |
| Positive Affect*               | .61 (.05) | .36 (.07) |         |
| Negative Affect*               |          |         |         |
| **Within-Individual Effects**  |          |         |         |
| Positive Affect                | .44 (.04) | .33 (.11) |         |
| Negative Affect                |          |         |         |
| **Between-Individual Variance** | 0.31 | 0.08 | 0.17 |
| **Within-Individual Variance** | 0.27 | 0.11 | 0.16 |
| \( R^2_{\text{between}} \)    | 0.73 | 0.45 | 0.69 |
| \( R^2_{\text{within}} \)     | 0.58 | 0.41 | 0.62 |

Note. Standard errors are in parentheses. ICC = .53.
was .33 (95% CI [.11, .54], \( r_{\text{effect}} = .21, p = .003 \)) and the impact of negative affect was \(-.21\) (95% CI \([- .39, -.02], r_{\text{effect}} = -.15, p = .027 \)) together explaining 62% of the within-person variance in daily positive relationships.

Using a state/daily measure of positive relationships, there was substantial within-person variance, about 47%. In addition, there was a strong impact of affect for state positive relationships (\( R^2 = .62 \)).

Of note, the impact of affect found for these dimensions was lower than that found for state satisfaction (100%; E. Jayawickreme et al., 2017b).

**Discussion**

As scholars of well-being endeavor to examine patterns of momentary well-being states over time, it is important to investigate whether state-based measurement of well-being indexes dimensions of well-being that are separable from current affect. This investigation is especially important in view of previous within-person work suggesting that state affect accounts for 100% of the within-person variance in state life satisfaction, and that individuals may rely on current affect in making short-term or ‘in-the-moment’ appraisals of their current life circumstances (E. Jayawickreme et al., 2017a). The present results similarly indicated a substantial within-person relation between affect and all key dimensions of eudemonic well-being, with positive affect exhibiting a stronger relation to EWB than negative affect. However, variance in EWB dimensions explained by current affect ranged from 60% (meaning) to 73% (evaluation), suggesting that state assessments of EWB (meaning, core self-evaluations, authenticity and gratitude) do provide unique information over and above current affect. Given the current controversy over the extent to which SWB and EWB well-being are empirically distinguishable (e.g., Disabato et al., 2016; Kashdan et al., 2008), our results suggest an intermediate position. Although positive affect is in fact associated with state eudaimonia across multiple dimensions, the relationship does not appear so strong as to render these constructs interchangeable (unlike in the instance of state satisfaction).

One interpretation of these results is that the level of positive affect one feels may in fact be an appropriate source of relevant information, a phenomenon that is consistent with past results linking state positive affect with state meaning (King et al., 2006). However, people likely rely on other sources of information as well, and future research should identify these sources. For example, recent research has found that a sense of mattering (i.e., feeling that one’s actions make a difference in the world and that life is worth living) was a significant precursor to trait assessments of meaning (Costin & Vignoles, 2019). Examining the relationship between state assessments of mattering and state meaning would be an interesting question for future research. One possibility is that progress towards meaningful or self-transcendent goals may predict increased state meaning (Yeager et al., 2014). It is further possible that perceived proximity to close others is a situational contingency that predicts state assessments of positive relationships (Fleeson, 2007). Future research should examine specific situational contingencies associated with state changes in EWB.

We should note a number of limitations. For one, we defined state well-being in this study as attained well-being during the current day (following E. Jayawickreme et al., 2017b). However, substantial variability exists in range of the duration of what is considered a ‘momentary’ assessment in the psychological literature. As noted in E. Jayawickreme et al. (2017b, p. 30), King et al.’s (2006), ‘momentary’ assessments of PWB assessed participants’ reflections over the past two days. Blackie et al. (2017), on the other hand, assessed state EWB five times per day. Modifying the frame of state assessments of well-being may vary the degree of impact by affect. It is conceivable that affect and dimensions of EWB exhibit lower or greater within-person association within shorter increments of time. Future research should examine the question of the ideal time frame for assessing state EWB and take into account the possibility that momentary assessments of EWB may not exhibit the same variability as momentary assessments of SWB (see Blackie et al., 2017 for data on the variability of momentary assessments of EWB).

Second, our results were based on an online sample collected on Amazon.com’s Mechanical Turk platform and might not generalize to other samples (although see Buhrmester et al., 2011). Third, the data were collected at a single point each day, as opposed to random momentary assessments throughout the day. Thus, despite the benefit of the within-person approach outlined in this paper, the retrospective nature of the daily reports may render them biased in some manner. However, given that our focus was on well-being as opposed to specific life events, end-of-day reports may in fact be a more optimal strategy for individuals to evaluate their quality of life, and momentary assessments may not the ideal strategy for assessment state EWB (see above). Fourth, our data were observational. We assumed (and our model predicts) that changes in affect cause changes in EWB. However, we note that our observed effects may be inflated (and possibly even completely explained) by unobserved third-variable confounds. An inherent tradeoff unfortunately exists.
between internal and external validity, as isolating the impact of affect on satisfaction independent of the affect manipulation would be difficult, even if affect were randomly manipulated. Similar to our past work, we, therefore, focused on maximizing ecological validity, given the nature of our research question (E. Jayawickreme et al., 2017a; see also, 2017b, p. 30). Fifth, we did not preregister this study. These data were collected in 2014, before preregistering became more prevalent in psychological research.

In conclusion, the present study utilized multi-level structural equation modeling (ML-SEM) with affect as a time-varying covariate to examine the within-person relationship between current affect and state EWB. We found that the impact of affect on EWB was less consequential than for state satisfaction observed in prior research, and that the impact of positive affect on these dimensions was stronger than the impact of negative affect. This implies that while state affect and EWB are highly related, they are not interchangeable. Further research should examine other predictors of the eudaimonia that people experience in their daily lives.

Note
1. To estimate multilevel reliability (see Geldhof et al., 2014), we ran multilevel confirmatory factor models and computed omega using the following formula: \( \omega = \frac{(\sum \lambda)^2}{(\sum \lambda)^2 + \sum \psi^2} \).

Authors’ contributions
EJ, ET and LB conceived the study; EJ, LB and BW developed materials and collected data; ET ran the analyses; EJ and ET drafted the paper; all authors provided feedback and edits on the final document.

Availability of data and materials
All data and materials are available from the first author upon request.

Compliance with ethical standards
Ethical approval: All procedures performed in studies involving human participants were in accordance with the ethical standards of the institutional research committee and with the American Psychological Association’s ethical standards. The study was reviewed and approved by the Wake Forest IRB prior to data collection (IRB # IRB00021398). No study involved research on animals.

Disclosure statement
The authors have no known conflicts of interest to disclose.

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Code availability
All code utilized in the analyses are available from the first author upon request.

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