Interpreting Primal-Dual Algorithms for Constrained MARL

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Abstract

Constrained multiagent reinforcement learning (C-MARL) is gaining importance as MARL algorithms find new applications in real-world systems ranging from energy systems to drone swarms. Most C-MARL algorithms use a primal-dual approach to enforce constraints through a penalty function added to the reward. In this paper, we study the structural effects of this penalty term on the MARL problem. First, we show that the standard practice of using the constraint function as the penalty leads to a weak notion of safety. However, by making simple modifications to the penalty term, we can enforce meaningful probabilistic (chance and conditional value at risk) constraints. Second, we quantify the effect of the penalty term on the value function, uncovering an improved value estimation procedure. We use these insights to propose a constrained multiagent advantage actor critic (C-MAA2C) algorithm. Simulations in a simple constrained multiagent environment affirm that our reinterpretation of the primal-dual method in terms of probabilistic constraints is effective, and that our proposed value estimate accelerates convergence to a safe joint policy.

Keywords: Multiagent reinforcement learning, primal-dual methods, chance constraints, conditional value at risk

1. Introduction

As reinforcement learning (RL) algorithms progress from virtual to cyber-physical applications, it will be necessary to address the challenges of safety, especially when systems are controlled by multiple agents. Examples of multiagent safety-critical systems include power grids Cui et al. (2022), building energy management (BEM) systems Biagioni et al. (2022), autonomous vehicle navigation Zhou et al. (2022), and drone swarms Chen et al. (2020). In each of these applications, agents must learn to operate in a complicated environment while satisfying various local and system-wide constraints. Such constraints, derived from domain-specific knowledge, are designed to prevent damage to equipment, humans, or infrastructure or to preclude failure to complete some task or objective.

Constrained multiagent reinforcement learning (C-MARL) poses challenges beyond the single-agent constrained reinforcement learning (C-RL) problem because the interactions between agents can influence both the satisfaction of constraints and the convergence of policies. The potential scale of C-MARL problems eliminates the possibility of directly using common model-based methods for C-RL, such as in Chen et al. (2018); Ma et al. (2021); Tabas and Zhang (2022). The main strategy for tackling C-MARL problems found in the literature is the Lagrangian or primal-dual method

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(see, e.g. Lu et al. (2021); Li et al. (2020); Lee et al. (2018); Parnika et al. (2021) and the references therein). Our aim is to understand some potential drawbacks of this approach and how they can be mitigated.

In the primal-dual approach to C-MARL, each agent receives a reward signal that is augmented with a penalty term designed to incentivize constraint satisfaction. The magnitude of the penalty term is tuned to steer policies away from constraint violations while not unnecessarily overshadowing the original reward. Although this approach has been shown to converge to a safe joint policy under certain assumptions Lu et al. (2021), it changes the structure of the problem in a way that is not well understood, leading to two challenges.

The first challenge is that the primal-dual algorithm only enforces discounted sum constraints derived from the original safety constraints of the system. As we will show, discounted sum constraints guarantee safety only in expectation, which is difficult to interpret. We propose simple modifications to the penalty term that enable the enforcement of more interpretable constraints, namely, chance constraints Mesbah (2016) and conditional value at risk constraints Rockafellar and Uryasev (2000), providing bounds on the probability and the severity of future constraint violations. There have been several C-RL algorithms that work with risk sensitivities García and Fernández (2015); Chow et al. (2018), but the multiagent context is less studied, and our contributions provide a novel understanding of the safety guarantees provided by C-MARL algorithms.

The second challenge is the fact that the reward is constantly changing as the dual variables are updated, which diminishes the accuracy of value estimates. We quantify this loss of accuracy, and we propose a new value estimation procedure to overcome it. Our proposal builds on results in Tessler et al. (2019) showing the linearity of the value function in the dual variables. We develop a novel class of temporal difference algorithms for value function estimation that directly exploits this observation, giving rise to a value estimate that maintains an accurate derivative with respect to the dual variables. Compared to existing algorithms, our estimates are much more robust to dual variable updates.

The specific C-MARL formulation we study in this paper is inspired by the BEM problem Molina-Solana et al. (2017); Biagioni et al. (2022), illustrated in Figure 1. The main objective of BEM is to control a building’s resources to minimize the cost of energy consumption while maintaining comfort and convenience for the occupants. However, when BEMs are deployed in multiple buildings, it is critical to ensure that the power network connecting them is safely operated because the uncoordinated control of buildings can cause network-level voltage or power constraints to be violated. This mandates a level of coordination among agents in the learning stage; thus, we adopt the commonly-studied centralized training/decentralized execution (CTDE) framework Lowe et al. (2017); Foerster et al. (2018), in which a simulator or coordinator provides global state information, constraint evaluations, and Lagrange multipliers (dual variables) to each agent during training. During the testing
The natural numbers and the real numbers are denoted \(\mathbb{N}\) and \(\mathbb{R}\), respectively. Given a measurable set \(S\), the set of all possible probability densities over \(S\) is denoted as \(\Delta_S\). For any discount factor \(\gamma \in (0,1)\) and any sequence \(\{y_t\}_{t=0}^\infty\), the discounted sum operator is \(\Gamma_T^{y_0} \equiv \sum_{t=0}^T \gamma^t y_t\), and \(\Gamma_\infty^{y_0} \equiv \lim_{T \to \infty} \Gamma_T^{y_0}\) if the limit exists. We often drop the second argument \(\gamma\) for brevity. The positive component operator is \([y]+ = \max\{y, 0\}\), and the logical indicator function \(I[\cdot]\) maps \{True, False\} to \{1, 0\}.

### 1.1. Notation

The setting is described by the tuple \(\left\{\mathcal{X}_i\right\}_{i \in \mathcal{N}}, \left\{\mathcal{U}_i\right\}_{i \in \mathcal{N}}, \{R_i\}_{i \in \mathcal{N}}, f, C, p_0, \gamma\), where \(\mathcal{N}\) is the index set of agents, \(\mathcal{X}_i \subset \mathbb{R}^n_i\) and \(\mathcal{U}_i \subset \mathbb{R}^m_i\) are the state and action spaces of agent \(i\), and \(R_i : \mathcal{X}_i \times \mathcal{U}_i \to \mathbb{R}\) is the reward function of agent \(i\). We assume that the sets \(\mathcal{X}_i\) and \(\mathcal{U}_i\) are compact for all \(i\). Let \(\mathcal{X} = \prod_{i \in \mathcal{N}} \mathcal{X}_i\) and \(\mathcal{U} = \prod_{i \in \mathcal{N}} \mathcal{U}_i\) be the joint state and action spaces of the system, respectively. Then \(f : \mathcal{X} \times \mathcal{U} \to \Delta_\mathcal{X}\) describes the state transition probabilities, i.e., \(f(\cdot \mid x, u)\) is a probability density function. The function \(C : \mathcal{X} \to \mathbb{R}^m\) is used to describe a set of safe states, \(S = \{x \in \mathcal{X} \mid C(x) \leq 0\}\).

Let \(p_0 \in \Delta_\mathcal{X}\) denote the initial state probability density and \(\gamma \in (0,1)\) be a discount factor. At time \(t\), the state, action, and reward of agent \(i\) are \(x_i^t, u_i^t, \) and \(r_i^t\), respectively, and constraint \(j\) evaluates to \(c_i^t = C_j(x_i^t)\). Using a quantity without a superscript to represent a stacked vector ranging over all \(i \in \mathcal{N}\) or all \(j \in \{1, \ldots, m\}\), a system trajectory is denoted \(\tau = \{(x_t, u_t, r_t, c_t)\}_{t=0}^\infty\).

In the noncooperative C-MARL framework, each agent seeks to learn a policy \(\pi_i : \mathcal{X}_i \to \Delta_\mathcal{U}_i\) that maximizes the expected discounted accumulation of individual rewards. We let \(\pi : \mathcal{X} \to \Delta_\mathcal{U}\) denote the joint policy, and \(f_\pi : \mathcal{X} \to \Delta_\mathcal{X}\) is the state transition probability induced by a joint policy \(\pi\). The tuple \((p_0, f, \pi)\) induces a state visitation probability density at each time step, \(p_t(x) = \int_{\mathcal{X}_t} p_0(x_0) \cdot \prod_{k=1}^t f_\pi(x_k \mid x_{k-1}) \, dx_0 \ldots \, dx_{k-1}\), and we say \(p_\infty(x) = \lim_{t \to \infty} p_t(x)\) for each \(x \in \mathcal{X}\) if the limit exists. The collection of visitation probabilities \(\{p_t\}_{t=0}^\infty\) gives rise to

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1. Even in buildings with advanced metering infrastructure or smart meters, they typically only exchange information with the utility a few times a day.
a probability density of trajectories $\tau$, denoted $\mathcal{M} \in \Delta [\mathbb{P}_{\infty}^{\mathbb{R}_n \times \mathbb{R}_m}]$; thus, the objective of each agent can be stated precisely as maximizing $\mathbb{E}_{\tau \sim \mathcal{M}}[\Gamma^{\infty}_{t=0} r^i_t]$.

The agents, however, must settle on a joint policy that keeps the system in the safe set $\mathcal{S}$. Due to the stochastic nature of the system, satisfying this constraint at all times is too difficult and in some cases too conservative. A common relaxation procedure is to formulate an augmented reward $\tilde{r}^i_t = r^i_t - \lambda^T c_t$ where $\lambda \in \mathbb{R}^m_+$, the Lagrange multiplier or dual variable, is adjusted to incentivize constraint satisfaction. This leads to the primal-dual algorithm for C-MARL, discussed in the next section. The following mild assumption facilitates the analysis.

**Assumption 1** $R^i, C^j, \text{and } p_t$ are bounded on $\mathcal{X}$ for all $i \in \mathcal{N}, \text{ all } j \in \{1, \ldots, m\}, \text{ and all } t \in \mathbb{N}$.

The boundedness of $R^i$ and $C^j$ is a common assumption Lu et al. (2021); Tessler et al. (2019); Paternain et al. (2019) that we will use to exchange the order of limits, sums, and integrals using the dominated convergence theorem. The assumption of bounded $p_t$ is not strictly necessary, but we use it throughout the paper to apply the dominated convergence theorem using the standard Lebesgue measure.

### 2.2. Primal-dual algorithms

The augmented reward function leads to the following min-max optimization problem for agent $i$:

$$
\min \lambda \geq 0 \max_{\pi_i} \mathbb{E}_{\tau \sim \mathcal{M}} [\Gamma^{\infty}_{t=0} [r^i_t - \lambda^T c_t]]
$$

$$
= \min \lambda \geq 0 \max_{\pi_i} \left( \mathbb{E}_{\tau \sim \mathcal{M}} [\Gamma^{\infty}_{t=0} [r^i_t]] - \lambda^T \mathbb{E}_{\tau \sim \mathcal{M}} [\Gamma^{\infty}_{t=0} [c_t]] \right)
$$

where (2) uses absolute convergence (stemming from Assumption 1) to rearrange the terms of the infinite sum. Note that the minimization over $\lambda$ is coupled across agents. Any fixed point of (2) will satisfy $\mathbb{E}_{\tau \sim \mathcal{M}} [\Gamma^{\infty}_{t=0} c_t] \leq 0$ because if $\mathbb{E}_{\tau \sim \mathcal{M}} [\Gamma^{\infty}_{t=0} c^j_t] \neq 0$, then the objective value can be reduced by increasing or decreasing $\lambda_j$, unless $\mathbb{E}_{\tau \sim \mathcal{M}} [\Gamma^{\infty}_{t=0} c^j_t] < 0$ and $\lambda_j = 0$. In other words, the primal-dual method enforces a discounted sum constraint derived from the safe set $\mathcal{S}$. Although discounted sum constraints are convenient, it is not obvious what they imply about safety guarantees with respect to the original constraints. We begin our investigation of discounted sum constraints by taking a closer look at a state visitation probability density known as the occupation measure.

### 3. Occupation measure

**Definition 2** (Occupation measure Paternain et al. (2019)) The occupation measure $\mu_\gamma \in \Delta_X$ induced by a joint policy $\pi$ is defined for any $x \in \mathcal{X}$ as $\mu_\gamma(x) = \Gamma^{\infty}_{t=0} p_t(x)$.

In this section, we provide some interpretations for the occupation measure before using it to ascribe meaning to discounted sum constraints. The first question one might ask is whether $\mu_\gamma$ is itself a pdf. It is, of course, nonnegative, and the following proposition shows it integrates to unity under mild conditions.

**Proposition 3** Under Assumption 1, $\int_X \mu_\gamma(x) dx = 1$. 
The proof for Proposition 3 is in Appendix A. What does $\mu_\gamma$ tell us about the behavior of a system under a given policy? It describes the probability of visiting a certain state but with more weight placed on states that are likely to be visited earlier in time. In fact, $\mu_\gamma$ describes the near-term behavior in the following sense.

**Proposition 4** Under Assumption 1, for any $x \in X$, the following statements hold:

1. $\lim_{\gamma \to 0^+} \mu_\gamma(x) = p_0(x)$.
2. $\lim_{\gamma \to 1^-} \mu_\gamma(x) = \lim_{t \to \infty} p_t$ if the latter limit exists.

The proof for Proposition 4 is in Appendix A. Figure 2 provides an illustration of the result in Proposition 4 when $p_t$ evolves as a normal distribution with mean $0.95t$ and constant variance. The point at which $\mu_\gamma$ equally resembles $p_0$ and $p_\infty$ is exactly at $\gamma = 0.95$.

According to Proposition 4, the occupation measure describes a state distribution that lies between the initial and long-term behavior of the system. But where exactly does it lie in between these two extremes? The effective horizon of a discounted planning problem is often set to $T_1(\gamma) = \frac{1}{1-\gamma}$, which is the expected termination time if the probability of an episode terminating at any given time step is $(1-\gamma)$ Paternain et al. (2022); however, the concept of a random stopping time might not be sensible in all applications. Another way to define the effective horizon is to study the geometric accumulation of weights. In this case, the effective horizon can be measured as $T_2(\gamma, \varepsilon) = \min\{K \in \mathbb{N} : \Gamma^K_{t=0} [1] \geq 1 - \varepsilon\}$, where $\varepsilon \in (0, 1)$ is a tolerance. Using either of these two definitions, the occupation measure can be said to describe the behavior of the system up to the effective horizon. Specifically, one may truncate the sum in Definition 2 at the effective horizon to obtain a conceptual understanding of what the occupation measure describes.

Depending on the application, either $T_1$ or $T_2$ can provide a more sensible connection between discounted and finite-horizon problems. But are these two definitions related? The next proposition answers this affirmatively by showing that $T_1$ is actually a special case of $T_2$.

**Proposition 5** $T_1(\gamma) = T_2(\gamma, \varepsilon)$ when $\varepsilon$ is set to $\frac{1}{\varepsilon} \approx 1$.

The proof for Proposition 5 is in Appendix A. Proposition 5 is illustrated in Figure 3, where the effective horizon is plotted as a function of $\gamma$ for three different values of $\varepsilon$. With an understanding of the occupation measure as a visitation density describing behavior up to the effective horizon, we can begin to derive meaningful risk-related interpretations of discounted sum constraints. These interpretations lead directly to sensible recommendations for the design of C-MARL algorithms.

4. Discounted risk metrics

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2. The full paper with appendix is available at [https://arxiv.org/abs/2211.16069](https://arxiv.org/abs/2211.16069).
The discounted sum constraint can naturally be reinterpreted as a certain type of average constraint. In particular, Assumption 1 ensures the equivalence $\mathbb{E}_{\tau \sim \mathcal{M}}[\Gamma_{t=0}^{\infty} C(x_t)] = \mathbb{E}_{x \sim \mu_x}[C(x)]$ [Paternain et al. (2019)]. This near-term average does not relate to any well-known risk metrics and hence does not provide a practical safety guarantee. In general, information about the mean of a distribution cannot be used to infer information about its tails; however, simple changes to the penalty function can yield information about either the probability of incurring a constraint violation or the expected severity of constraint violations.

**Proposition 6 (Near-term probability of constraint violations)** Suppose that for some $\delta_j \in [0, 1]$ and $\alpha_j \in \mathbb{R}$, we have $\mathbb{E}_{\tau \sim \mathcal{M}}[\Gamma_{t=0}^{\infty} I[C^j(x_t) \geq \alpha_j]] \leq \delta_j$. Then under Assumption 1, $\Pr\{C^j(x) \geq \alpha_j \mid x \sim \mu_x\} \leq \delta_j$.

**Proof** $\mathbb{E}_{\tau \sim \mathcal{M}}[\Gamma_{t=0}^{\infty} I[C^j(x_t) \geq \alpha_j]] = \mathbb{E}_{x \sim \mu_x}[I[C^j(x) \geq \alpha_j]] = \Pr\{C^j(x) \geq \alpha_j \mid x \sim \mu_x\}$. The first equality uses Assumption 1 to apply an equivalence established in e.g. Paternain et al. (2019). The second equality follows from the definition of expectation.

Proposition 6 makes it easy to enforce chance constraints using primal-dual methods. When the penalty term $C^j(x)$ is replaced by the quantity $I[C^j(x) \geq \alpha_j] - \delta_j$, the primal-dual algorithm enforces $\mathbb{E}_{\tau \sim \mathcal{M}}[\Gamma_{t=0}^{\infty} I[C^j(x_t) \geq \alpha_j]] - \delta_j \leq 0$. By Proposition 6, this guarantees that $\Pr\{C^j(x) \geq \alpha_j \mid x \sim \mu_x\} \leq \delta_j$. Because the probability of constraint violations is defined with $x$ varying over $\mu_x$, we call the resulting guarantee a near-term or discounted chance constraint. This can be repeated for each $j \in \{1, \ldots, m\}$, providing a set of bounds on the probability of violating each constraint by more than its tolerance $\alpha_j$. On the other hand, we can control the probability of violating any constraint as follows. Define the statement $C(x) \geq \alpha$ to be true if $C^j(x) \geq \alpha_j \forall j \in \{1, \ldots, m\}$, and false otherwise. Then, applying Proposition 6 to the test condition $C(x) \geq \alpha$ will result in a bound on $\Pr\{C(x) \geq \alpha \mid x \sim \mu_x\}$.

While discounted chance constraints enable one to control the probability of extreme events in the near future, conditional value at risk constraints Rockafellar and Uryasev (2000) afford control over the severity of such events.

**Definition 7 (Rockafellar and Uryasev (2000))** Given a risk level $\beta \in (0, 1)$, a cost $h : \mathcal{X} \rightarrow \mathbb{R}$, and a probability density $\mu$ on $\mathcal{X}$, the value at risk (VaR) and conditional value at risk (CVaR) are defined as:

$$\text{VaR}(\beta, h, \mu) = \min\{\alpha \in \mathbb{R} : \Pr\{h(x) \leq \alpha \mid x \sim \mu\} \geq \beta\},$$

$$\text{CVaR}(\beta, h, \mu) = \frac{1}{1 - \beta} \int_{h(x) \geq \text{VaR}(\beta, h, \mu)} h(x)\mu(x)dx.$$

In other words, $\text{VaR}(\beta, h, \mu)$ is the least upper bound on $h$ that can be satisfied with probability $\beta$, while $\text{CVaR}(\beta, h, \mu)$ describes the expected value in the VaR-tail of the distribution of $h$. CVaR characterizes the expected severity of extreme events, which can be defined precisely as the $(1 - \beta)$
fraction of events \( x \) with the worst outcomes as ranked by the cost incurred, \( h(x) \). The VaR and CVaR for \( h(x) = x \), when \( x \) follows a standard normal distribution, are illustrated in Figure 4, where the shaded region has an area of \( (1 - \beta) \). For the rest of the paper, we assume that the cdf of \( h(x) \) is continuous when \( x \sim \mu \). For further details and for cases in which this assumption does not hold, we refer the reader to Rockafellar and Uryasev (2002).

**Proposition 8 (Near-term CVaR)** For any \( \alpha_j \geq 0 \), suppose that \( \mathbb{E}_{\tau \sim \mathcal{M}}[\Gamma_{t=0}^{\infty}[(C^j(x_t) - \alpha_j)_+]] \leq \delta_j \). Then, CVaR\((\beta, C^j, \mu, \gamma)\) \( \leq \alpha_j + (1 - \beta)^{-1}\delta_j \).

**Proof** Under Assumption 1, the identity \( \mathbb{E}_{\tau \sim \mathcal{M}}[\Gamma_{t=0}^{\infty}[(C^j(x_t) - \alpha_j)_+]] = \mathbb{E}_{x \sim \mu, \gamma}[(C^j(x) - \alpha_j)_+] \) holds Paternain et al. (2019). Next, we use the fact that the CVaR is the minimum value of the convex dual algorithm. Here, the penalty term used is \( \frac{1}{\beta} \). For further details and for cases in which this assumption does not hold, we refer the reader to Rockafellar and Uryasev (2000); thus, \( F \) provides an upper bound on CVaR. Some rearranging leads to the result.

![Figure 4: Example of VaR and CVaR at risk level \( \beta = 0.9 \).](image)

Similar to the chance-constrained case, Proposition 8 makes it easy to enforce CVaR constraints in the primal-dual algorithm. Here, the penalty term used is \( [C^j(x) - \alpha_j]_+ - \delta_j \). Using this penalty, the algorithm enforces \( \mathbb{E}_{\tau \sim \mathcal{M}}[\Gamma_{t=0}^{\infty}[(C^j(x_t) - \alpha_j)_+]] - \delta_j \leq 0 \), which by Proposition 8 implies CVaR\((\beta, C^j, \mu)\) \( \leq \alpha_j + (1 - \beta)^{-1}\delta_j \). By repeating for each \( j \in \{1, \ldots, m\} \), we can bound the expected severity of the constraint violations for each of the \( m \) constraints. Because the CVaR constraint is defined with \( x \) varying over \( \mu, \gamma \), the resulting guarantee is called a near-term or discounted CVaR constraint.

To obtain a tight bound on the CVaR, \( \alpha_j \) must be set to VaR\((\beta, C^j, \mu)\), which minimizes the function \( F \) introduced in the proof of Proposition 8 Rockafellar and Uryasev (2000). Unfortunately, the VaR is not known ahead of time. Chow et al. (2018) include \( \alpha_j \) as an optimization variable in the learning procedure, but extending their technique to the multiagent setting is not straightforward. Our approach is to include it as a tunable hyperparameter. Simulation results in Section 6 show that it is easy to choose \( \alpha_j \) to give a nearly tight bound.

### 5. Primal-dual value functions

In this section, we investigate challenges with value estimation in the primal-dual regime. The fact that the reward to each agent is constantly changing (due to dual variable updates) makes it difficult to accurately estimate state values. To quantify this decrease in accuracy, we introduce the value functions induced by the joint policy \( \pi \), \( \{V_{\pi}^i : \mathcal{X} \times \mathbb{R} \to \mathbb{R} \}_{i \in \mathcal{N}} \); \( \{V_{R,\pi}^i : \mathcal{X} \to \mathbb{R} \}_{i \in \mathcal{N}} \); \( V_{C,\pi}^i : \mathcal{X} \to \mathbb{R}^m \) where:

\[
V_{\pi}^i(x, \lambda) = \mathbb{E}_{\tau \sim \mathcal{M}}[\sum_{t=0}^{\infty} r_t^i - \lambda^T c_t | x_0 = x],
\]

\[
V_{R,\pi}^i(x) = \mathbb{E}_{\tau \sim \mathcal{M}}[\sum_{t=0}^{\infty} r_t^i | x_0 = x], \quad V_{C,\pi}^i(x) = \mathbb{E}_{\tau \sim \mathcal{M}}[\sum_{t=0}^{\infty} \lambda^T c_t | x_0 = x].
\]

Note that \( c_t \) could be modified as indicated in Section 4, and the following results would hold for the modified penalty function.
Obviously, it is impossible to learn an accurate value function when \( \lambda \) is unknown and changing; however, simply making \( \lambda \) available to a value function approximator does not guarantee good generalization beyond previously seen values of \( \lambda \). Having a good estimate of the derivative of the value function with respect to \( \lambda \) will ensure accuracy under small perturbations to the dual variables. Fortunately, this derivative is easy to obtain. Formally, under Assumption 1, we can write \( V_\lambda^i(x,\lambda) = V_{R,\pi}^i(x) - \lambda^T V_{C,\pi}(x) \) Tessler et al. (2019), and therefore, \( \nabla_\lambda V_\lambda^i(x,\lambda) = -V_{C,\pi}(x) \).

By learning \( V_{R,\pi}^i \) and \( V_{C,\pi} \) as separate functions and then combining them using the true value of \( \lambda \), we can construct a value estimate whose derivative with respect to the dual variables is as accurate as our estimate of \( V_{C,\pi} \) itself. This estimate will be more robust to small changes in \( \lambda \). We will refer to this type of value estimate as a structured value function or a structured critic.

**Proposition 9** Let \( \tilde{c} = \mathbb{E}_{x \sim \mu_x} [C(x)] \) and \( \Sigma^2_C = \mathbb{E}_{x \sim \mu_x} [(\tilde{c} - C(x))(\tilde{c} - C(x))^T] \). Suppose \( \lambda \) is randomly varying with mean \( \bar{\lambda} \) and variance \( \Sigma^2_\lambda \). Using a structured value function approximator can reduce the mean square temporal difference error by up to \( \text{Tr}[(\Sigma^2_\lambda \cdot (\Sigma^2_C + \tilde{c} \tilde{c}^T))] \).

The proof of Proposition 9 is in Appendix A. Figure 5 illustrates Proposition 9 in a simple value estimation task with quadratic rewards, linear dynamics and policies, linear state constraints, and randomly varying \( \lambda \). The *generic critic* (GC) is a value function modeled as a quadratic function of the state only. The *input-augmented critic* (IAC) is a value function modeled as an unknown quadratic function of the state and dual variables, while the *structured critic* (SC) is modeled using \( \hat{V}^i_\pi = \hat{V}^i_{R,\pi} - \lambda^T \hat{V}_{C,\pi} \) with quadratic \( \hat{V}^i_{R,\pi} \) and linear \( \hat{V}_{C,\pi} \) trained on their respective signals.

The dashed line in Figure 5 is at the value \( \text{Tr}[\Sigma^2_\lambda \cdot (\Sigma^2_C + \tilde{c} \tilde{c}^T)] \) predicted in Proposition 9. In this simple value estimation task, high accuracy can be achieved when conditioning on the randomly varying \( \lambda \); however, having an accurate estimate of \( \nabla_\lambda V_{\pi} \) by using a structured critic is also shown to help. Although the assumptions of Proposition 9 do not hold in general, the results in Section 6 show that using the structured value function still yields improved results.

The loss function for value function approximation is therefore given by:

\[
TDE(x, x') = [R^i(x^i) + \gamma \hat{V}^i_{R,\pi}((x^i)') - \hat{V}^i_{R,\pi}(x^i)]^2 + \sum_{j=1}^{m} [C^j(x) + \gamma \hat{V}^j_{C,\pi}(x') - \hat{V}^j_{C,\pi}(x)]^2
\]

where \( x \in \mathcal{X} \) and \( x' \sim f^\pi(x) \). Equation (5) is simply a sum of squared temporal difference errors over the set of \( m + 1 \) value functions. For algorithmic details, we refer the reader to Appendix B.

### 6. Simulations

In our simulations, we sought to demonstrate the effectiveness of the penalty modifications and structured critic proposed in sections 4 and 5. We tested our findings in a modified multiagent particle environment\(^3\) Lowe et al. (2017) with two agents pursuing individual objectives subject

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3. Code for the environments is available at [github.com/dtabas/multiagent-particle-envs](github.com/dtabas/multiagent-particle-envs).
to a constraint on the joint state. The state of each agent is its position and velocity in \( \mathbb{R}^2 \), i.e. \( x^i = [y^i]^T \), where \( y^i \in \mathbb{R}^2 \) is the position and \( v^i \in \mathbb{R}^2 \) is the velocity of agent \( i \). The objective of each agent is to drive its position \( y^i \) to a landmark \( y^{i*} \in \mathbb{R}^2 \), while making sure that the agent ensemble satisfies the safety constraint. The reward and constraint functions are given by:

\[
R^i(y^i) = -\xi_i \| y^i - y^{i*} \|_2^2, \quad C(y) = 1^T y
\]

where \( \xi_i > 0 \) is a constant and \( y = [y^1]^T \) is the position of the agent ensemble.

The landmark \( y^* = [y^{1*}^T \ y^{2*}^T]^T \) is stationed outside of the safe region \( S = \{ y \mid C(y) \leq 0 \} \). Thus, the agents cannot both reach their goals while satisfying \( C(y) \leq 0 \). To train the agents to interact in this environment, we used a modified version of the EPYMARL codebase\(^4\) Papoudakis et al. (2020). We tested several MARL algorithms, including MADDPG Lowe et al. (2017), COMA Foerster et al. (2018), and MAA2C Papoudakis et al. (2020). We decided to use the MAA2C algorithm because it consistently produced the best results and because as a value function-based algorithm, it provided the most straightforward route to implementing the changes proposed in Section 5. Details of the algorithm, pseudocode, hyperparameters, and supplementary simulation results are provided in Appendix B.

For each risk metric described in Section 4, we tested the convergence of the agents to a safe policy with and without modifications to the penalty and value functions. Figure 6 shows the results when we make the substitution \( C(x) \leftarrow I[C(x) \geq \alpha] - \delta \) in the penalty function to enforce a chance constraint, \( \text{Pr}\{C(x) \geq \alpha \mid x \sim \mu_\gamma\} \leq \delta \) with \( \alpha \) and \( \delta \) each set to 0.1. The modified penalty function performs the best as a chance constraint-enforcing signal (red and green lines in Figure 6). Whether or not the penalty function is modified, the structured critic finds safer policies throughout training (red vs. green and orange vs. blue lines).

Figure 7 shows the results when we make the substitution \( C(x) \leftarrow [C(x) - \alpha]^+ - \delta \) in the penalty function to enforce the constraint \( \text{CVaR}(\beta, C, \mu_\gamma) \leq \alpha + (1 - \beta)^{-1} \delta \). Using the modified penalty (red and green lines in Figure 7) drives the CVaR upper bound (drawn in dashed lines) to the target value, and due to the choice of \( \alpha \), this bound is nearly tight. On the other hand, using the original penalty results in an overly conservative policy that achieves low risk at the expense of rewards (right panel). We also point out that when using the modified penalty with the structured critic, the CVaR is lower throughout training compared to when the generic critic is used, indicating improved effectiveness in enforcing limits on risk.

We chose \( \alpha \) using the following heuristic, to make the bound on CVaR nearly tight. The “correct” value of \( \alpha \) that would achieve a tight bound is \( \text{VaR}(\beta, C, \mu_\gamma) \). Moreover, the upper bound that we used is convex and continuously differentiable in \( \alpha \) Rockafellar and Uryasev (2000); therefore, small errors in \( \alpha \) will lead to small errors in the upper bound on CVaR, and any approximation of

\(^4\) Code for the algorithms is available at [github.com/dtabas/epymarl](https://github.com/dtabas/epymarl).
7. Conclusion

In this paper, we studied the effect of primal-dual algorithms on the structure of C-MARL problems. First, we used the occupation measure to study the effect of the penalty term on safety. We showed that using the constraint function as the penalty enforces safety only in expectation, but by making simple modifications to the penalty term, one may enforce meaningful probabilistic safety guarantees, namely, chance and CVaR constraints. These risk metrics are defined over the occupation measure, leading to notions of safety in the near term. Next, we studied the effect of the penalty term on the value function. When the dual variable and constraint evaluation signals are available, it is easy to model the relationship between the penalty term and the value function. By exploiting this structure, the accuracy of the value function can be improved. We demonstrated the usefulness of both of these insights in a constrained multiagent particle environment, showing that convergence to a low-risk policy is accelerated. One open question is the effect of primal-dual methods on game outcomes. Some agents might pay a higher price than others for modifying their policies to satisfy system-wide constraints. Understanding and mitigating this phenomenon will be the focus of future work.
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Appendix A. Theoretical results

A.1. Proof of Proposition 3

Applying the definition of $\mu_\gamma$, we have $\int_X \mu_\gamma(x) dx = \int_X \Gamma_{t=0}^{\infty} p_t(x) dx$. Using the Dominated Convergence Theorem, we can exchange the order of the sum and integral. Each individual $p_t$ integrates to 1. The geometric sum property ensures that the resulting expression evaluates to 1.

A.2. Proof of Proposition 4

1. By definition, we have $\lim_{t \to 0^+} \mu_\gamma(x) = \lim_{t \to 0^+} \Gamma_{t=0}^{\infty} p_t(x)$. Using Tannery’s theorem, we can exchange the order of the limit and the infinite sum. Each zeroth term in the sum evaluates to $p_0(x)$ and all other terms evaluate to 0.

2. Assume $\lim_{t \to \infty} p_t$ exists, and denote it $p_\infty$. Using the triangle inequality, we have

$$|\mu_\gamma(x) - p_\infty(x)| \leq \sum_{t=0}^{\infty} |p_t^\theta(x) - p_\infty^\theta(x)|$$

for some $N \in \mathbb{N}$. Since $p_t(x) \to p_\infty(x)$, we can choose $N$ large enough to make the second term in (8) arbitrarily small. Then, using boundedness of $p_t$ for all $t$, we can take $\gamma \to 1^-$ to make the first term arbitrarily small.

A.3. Proof of Proposition 5

By the geometric sum property, we have $T_{wd}(\gamma, \varepsilon) = \min\{K \in \mathbb{N} : \sum_{t=0}^{K-1} 1 \geq 1 - \varepsilon\} = \min\{K \in \mathbb{N} : 1 - \gamma^K \geq 1 - \varepsilon\} = \min\{K \in \mathbb{N} : K \geq \frac{\log \varepsilon}{\log \gamma}\} = \left\lceil \frac{\log \varepsilon}{\log \gamma} \right\rceil$. Setting $T_{wd}(\gamma, \varepsilon) = T_{t}(\gamma)$ and solving for $\varepsilon$ (ignoring the integer constraint) yields $\varepsilon = \gamma^{1/\gamma}$. Finally, taking $\lim_{\gamma \to 1^-} \gamma^{1/\gamma}$ yields $1/\gamma$.

A.4. Proof of Proposition 9

Let $x \sim \mu_\gamma$, $x' \sim f^\pi(x)$, and $\Sigma_\gamma^2 = \mathbb{E}_{x \sim \mu_\gamma}(\ell - C(x)(\ell - C(x))^T)$. Suppose $\lambda$ is randomly distributed with mean $\bar{\lambda}$ and variance $\Sigma_\lambda^2$. For any value function approximator $\hat{V}_\pi^i$, assume $\lambda$ and $\hat{V}_\pi^i$ are independent. Let $\eta = [1 \ \lambda]^T$, $d = [R^i(x) \ C(x)^T]^T$, $\hat{V}_\pi^i : \mathcal{X} \to \mathbb{R}$, $\hat{V}_{R^i,\pi} : \mathcal{X} \to \mathbb{R}$, and $\hat{V}_{C,\pi} : \mathcal{X} \to \mathbb{R}^m$. Let $\mathcal{D}$ be a dataset of trajectories sampled from $\mathcal{M}$ that is used to train $\hat{V}_\pi^i$, $\hat{V}_{R^i,\pi}$, and $\hat{V}_{C,\pi}$. The mean square temporal difference error achieved by using a generic value function is

$$MSTDE_1 = \mathbb{E}_{x,x',\lambda,\mathcal{D}}[(\eta^T d + \gamma \hat{V}_\pi^i(x') - \hat{V}_\pi^i(x))^2]$$

while the error achieved using the structured value function is

$$MSTDE_2 = \mathbb{E}_{x,x',\mathcal{D}}[(\eta^T d + \gamma [\hat{V}_{R,\pi}(x') - \lambda^T \hat{V}_{C,\pi}(x')] - [\hat{V}_{R^i,\pi}(x) - \lambda^T \hat{V}_{C,\pi}(x))^2].$$
Note that in (10) we do not take the expectation over \( \lambda \) since the dual variables are made available to this function approximator.

Begin with the states and dual variables fixed at \((\bar{x}, \bar{x}', \bar{\lambda})\). Let \( \hat{g}(\bar{x}, \bar{x}') = \left[ \hat{V}_{R,\pi}(\bar{x}) \ V_{C,\pi}(\bar{x}') \right]^T \) and \( \hat{h}(\bar{x}, \bar{x}') = \hat{V}_{R}(\bar{x}) - \gamma \hat{V}_{C}(\bar{x}') \). Then, suppressing the arguments \((\bar{x}, \bar{x}')\) and setting \( \bar{\eta} = [1 \ -\bar{\lambda}^T]^T \), we can write the squared temporal difference error at \((\bar{x}, \bar{x}', \bar{\lambda})\) as

\[
STDE_1(\bar{\eta}) = \mathbb{E}_d[(\bar{\eta}^T d - \hat{h})^2],
\]

\[
STDE_2(\bar{\eta}) = \mathbb{E}_d[(\bar{\eta}^T d - \eta^T \hat{g})^2].
\]

The loss function used to train \( \hat{V}_{R,\pi} \) and \( \hat{V}_{C,\pi} \) is

\[
\mathbb{E}_D[||d - \hat{g}||^2].
\]

Since \( d \) is a deterministic function of \( x \), (13) can be decomposed into bias and variance terms:

\[
\mathbb{E}_D[||d - \hat{g}||^2] = \mathbb{E}_D[\sum_{k=0}^{m} (d_k - \hat{g}_k)^2]
\]

\[
= \sum_{k=0}^{m} \mathbb{E}_D[(d_k - \hat{g}_k)^2]
\]

\[
= \sum_{k=0}^{m} [(d_k - \mathbb{E}_D \hat{g}_k)^2 + \mathbb{E}_D[(\hat{g} - \mathbb{E}_D \hat{g})^2]]
\]

\[
= \sum_{k=0}^{m} [b_k^2 + \sigma_k^2]
\]

\[
= \text{Tr}[bb^T + \Sigma^2]
\]

where \( k = 0 \) corresponds to the reward signal and \( k = 1, \ldots, m \) corresponds to the cost signals.

Following a similar line of reasoning, we can use (18) to rewrite (12) as

\[
STDE_2(\bar{\eta}) = \text{Tr}[(bb^T + \Sigma^2)(\bar{\eta} \bar{\eta}^T)].
\]

For the sake of argument, we assume that \( \hat{g} \) and \( \hat{h} \) achieve the same performance at \((x, x', \lambda)\), that is,

\[
STDE_1(\bar{\eta}) = STDE_2(\bar{\eta}) = \text{Tr}[(bb^T + \Sigma^2)(\bar{\eta} \bar{\eta}^T)]
\]

(20)

where \( \text{Tr}[(bb^T)(\bar{\eta} \bar{\eta}^T)] \) and \( \text{Tr}[\Sigma^2 \bar{\eta} \bar{\eta}^T] \) reflect the bias squared and variance terms, respectively. How do \( STDE_1 \) and \( STDE_2 \) change when \( \lambda \) is allowed to vary? Using the generic estimator, the noise in \( \lambda \) will introduce some amount of irreducible error into \( STDE_1 \). On the other hand, using \( \lambda = \bar{\lambda} + \Delta \lambda \) in our proposed estimator will change the bias and variance terms in \( STDE_2 \) while the irreducible error remains at zero (since there is no uncertainty when \( \Delta \lambda \) is known). Setting \( \Delta \eta = [0 \ -\Delta \lambda^T]^T \), the temporal difference errors at \((\bar{x}, \bar{x}', \bar{\lambda} + \Delta \lambda)\) are

\[
STDE_1(\bar{\eta} + \Delta \eta) = \text{Tr}[(bb^T + \Sigma^2)(\bar{\eta} \bar{\eta}^T)] + (\Delta \eta^T d)^2,
\]

\[
STDE_2(\bar{\eta} + \Delta \eta) = \text{Tr}[(bb^T + \Sigma^2)((\bar{\eta} + \Delta \eta)(\bar{\eta} + \Delta \eta)^T)].
\]

(21)

(22)
Taking the expectation over $\Delta \lambda$ which has a mean of zero and a variance of $\Sigma^2_\lambda$, and setting $\Sigma^2_\eta = \begin{bmatrix} 0 & 0 \\ 0 & \Sigma^2_\lambda \end{bmatrix}$, yields

$$E_{\Delta \lambda}[S T D E_1(\bar{\eta} + \Delta \eta) - S T D E_2(\bar{\eta} + \Delta \eta)] = \text{Tr}[\Sigma^2_\eta (dd^T - bb^T - \Sigma^2)]$$

(23)

$$= \text{Tr}[\Sigma^2_\lambda (cc^T - bb^T - \Sigma^2)]$$

(24)

where $\tilde{b} = (c - E_D\hat{g}_C)$, $\tilde{\Sigma}^2 = E_D[(\hat{g}_C - E_D\hat{g}_C)^2]$, and $\hat{g}_C = \hat{V}_{C,\pi}(x) - \gamma \hat{V}_{C,\pi}(x')$. Note that $E_D[\|c - \hat{g}_C\|^2] = \text{Tr}[bb^T + \Sigma^2]$. Taking $\tilde{b}, \tilde{\Sigma}^2 \to 0$ as the accuracy of $\hat{g}_C$ improves, (24) can be estimated as

$$\text{Tr}[\Sigma^2_\lambda cc^T].$$

(25)

Taking the expectation over $c \sim C(x), x \sim \mu_\gamma$ yields the final result.

### Appendix B. Simulation details

#### B.1. Algorithm

The Constrained Multiagent Advantage Actor Critic (C-MAA2C) algorithm is shown in Algorithm 1. The main differences from the basic MAA2C algorithm are the penalty modifications in lines 9 and 11, the use of vector-valued value functions $\hat{V}^i : \mathcal{X} \to \mathbb{R}^{m+1}$ (one per agent in the noncooperative setting), and the dual update.

There are two apparent differences between Algorithm 1 and the concepts described in the main text. The first is that Algorithm 1 uses n-step returns in the advantage function, whereas Section 5 only considers one-step returns. We resolve this discrepancy by revisiting the proof of Proposition 9. First, note that the coefficients $\eta$ can be factored out of the returns just like they are factored out of the rewards. Thus, the proof only requires slight modifications up to the last line, Equation (25). Using returns instead of rewards in (25) will lead to a different numerical result but the conclusion (justification for using a structured value function) will be the same.

The second apparent difference is the fact that Algorithm 1 considers finite-horizon episodic tasks, thus the primal-dual algorithm will enforce $E_{\tau \sim \mathcal{M}}[\sum^T_{t=0} c_t] \leq 0$. Due to the finite horizon, we cannot directly use the occupation measure to interpret the meaning of this constraint. However, we can define the occupation measure over a finite horizon as

$$\mu_{\gamma,T}(x) = \frac{1}{1 - \gamma^{T+1}} \sum^T_{t=0} p_t(x).$$

(26)

It is easy to show that $\mu_{\gamma,T}$ is nonnegative and integrates to unity over $\mathcal{X}$. We can use $\mu_{\gamma,T}$ in place of $\mu_\gamma$ everywhere in order to interpret discounted sum constraints and to generate probabilistic constraints in finite-horizon episodic tasks. The statements $E_{\tau \sim \mathcal{M}}[\sum^T_{t=0} c_t] \leq 0$, $E_{\tau \sim \mathcal{M}}[(1 - \gamma^{T+1})^{-1} \sum^T_{t=0} c_t] \leq 0$, and $E_{x \sim \mu_\gamma}(C(x)) \leq 0$ are equivalent. Note that the effective horizon discussed in Section 3 may be shorter than the horizon length $T$. 
Algorithm 1 C-MA2C with probabilistic safety guarantees and structured value functions

1: Input discount factor $\gamma$, learning rates $\zeta_\theta$, $\zeta_\omega$, $\zeta_\lambda$, n-step return horizon $\kappa$, tolerances $\alpha$ and $\delta$, multiplier limit $\lambda_{\text{max}}$, episode length $T$, number of episodes $K$, risk metric $\in \{\text{average, chance, CVaR}\}$
2: Initialize actor parameters $\{\theta^i\}_{i \in \mathcal{N}}$, critic parameters $\{\omega^i\}_{i \in \mathcal{N}}$, parameterized policies $\pi^i(\cdot | \theta^i) : \mathcal{X}_i \to \Delta_{U_i}$, parameterized value estimates $\hat{V}^i(\cdot | \omega^i) : \mathcal{X} \to \mathbb{R}^{m+1}$, dual variables $\lambda \in \mathbb{R}^m$, $\eta := [1 - \lambda T]^T$
3: for $k = 1, 2, \ldots, K$ do
4: Initialize $x_0 \sim p_0$ ▶ Run 1 episode
5: for $t = 0, 1, \ldots, T$ do
6: Sample $u^i_t \sim \pi^i(\cdot | x^i_t, \theta^i)$ for $i \in \mathcal{N}$
7: Receive $\{r^i_t\}_{i \in \mathcal{N}}$, $c_t, x_{t+1}$
8: if risk metric $= \text{chance}$ then
9: $c_t \leftarrow I[c_t \geq \alpha] - \delta$ ▶ Proposition 6
10: else if risk metric $= \text{CVaR}$ then
11: $c_t \leftarrow [c_t - \alpha]_+ - \delta$ ▶ Proposition 8
12: end if
13: Let $d^i_t = [r^i_t \ c^T_t]$ for $i \in \mathcal{N}$
14: end for
15: for $i \in \mathcal{N}$ do
16: for $t = 0, 1, \ldots, T$ do
17: $N = \min\{T, t + \kappa\}$
18: $D^i_t = \sum_{n=t}^{N-1} \gamma^{n-t} d^i_n + \gamma^{N-t} \hat{V}^i (x_N | \omega^i)$ ▶ Compute n-step returns
19: $A^i_t = \eta^T (D^i_t - \hat{V}^i (x_t | \omega^i))$ ▶ Compute advantages
20: end for
21: $\theta^i \leftarrow \theta^i + \zeta_\theta \sum_{t=0}^{T} A^i_t \nabla_{\theta^i} \log \pi^i (u^i_t | x^i_t, \theta^i)$ ▶ Actor update
22: $\omega^i \leftarrow \omega^i - \zeta_\omega \nabla_{\omega^i} \sum_{t=0}^{T} \|D^i_t - \hat{V}^i (x_t | \omega^i)\|_2^2$ ▶ Critic update
23: end for
24: $\lambda \leftarrow \lambda + \zeta_\lambda \Gamma^T_{t=0} c_t$ ▶ Dual update
25: $\lambda \leftarrow \min\{[\lambda]_+, \lambda_{\text{max}}\}$
26: $\eta \leftarrow [1 - \lambda T]^T$
27: end for

B.2. Hyperparameters
Simulation hyperparameters are listed in Table 1.
| Simulation                                      |       |
|-----------------------------------------------|-------|
| Episode length                                | 25    |
| Number of episodes                            | \{4, 8\} × 10^4 |
| Number of trials per configuration            | 5     |

| RL algorithm                                  |       |
|-----------------------------------------------|-------|
| Discount factor \( \gamma \)                 | 0.99  |
| Actor learning rate \( \zeta_\theta \)        | 3 × 10^{-4} |
| Critic learning rate \( \zeta_\omega \)       | 3 × 10^{-4} |
| Dual update step size \( \zeta_\lambda \)     | 1 × 10^{-4} |
| Optimizer                                     | Adam(\( \beta_{\text{Adam}} = (0.9, 0.999) \)) |
| n-step return horizon \( \kappa \)            | 5     |

| Constraint enforcement                        |       |
|-----------------------------------------------|-------|
| \( \lambda_{\text{max}} \)                   | 10    |
| Risk level \( \beta \)                       | 0.9   |
| “LHS tolerance” \( \alpha \):                |       |
| Average constraints                           | N/A   |
| Chance constraints                            | 0.1   |
| CVaR constraints                              | 0.2   |
| “RHS tolerance” \( \delta \):                |       |
| Average constraints                           | 0     |
| Chance constraints                            | 0.1   |
| CVaR constraints                              | 5 × 10^{-3} |

| Actors                                         |       |
|-----------------------------------------------|-------|
| Policy architecture                           | Multi-layer perceptron |
| Number of hidden layers                       | 2     |
| Hidden layer width                            | 64    |
| Hidden layer activation                       | ReLU  |
| Output layer activation                       | Linear|
| Action selection                              | Categorical sampling |
| Parameter sharing                             | No    |

| Critics                                        |       |
|-----------------------------------------------|-------|
| Critic architecture                           | Multi-layer perceptron |
| Number of hidden layers                       | 2     |
| Hidden layer width                            | 64    |
| Hidden layer activation                       | ReLU  |
| Output layer activation                       | Linear|
| Target network update interval                 | 200 episodes |
| Parameter sharing                             | No    |

Table 1: Simulation hyperparameters.
B.3. Additional simulation results

Here, we provide some additional results to supplement the findings in Section 6. First, we compared the convergence to a safe policy under the original discounted sum constraint and found that similar to the results for the other types of constraints, the structured critic demonstrates a better safety margin throughout training. This is illustrated in Figure 8.

![Figure 8: Evaluation of the discounted sum constraint throughout training, showing that the structured critic helps the actor to find safer policies faster. Each line and shaded region represents the mean and standard deviation over 5 training runs. Key: SC = structured critic.](image)

Next, we provide a closer look at the accuracy of the CVaR upper bound provided in Proposition 8, and illustrated using dashed lines in the left panel of Figure 7. Table 2 shows that in all four configurations in which the CVaR was evaluated, the upper bound is a fairly accurate estimate. The results from Section 6 show that this upper bound can be used to drive the actual CVaR below a target value. Although using a structured critic with modified penalty function yielded the most accurate CVaR UB, the accuracy in all four configurations could be improved by making further adjustments to the tolerance $\alpha$. The error is reported for policies tested at the end of the training phase.

| Penalty function | Critic     | CVaR UB error |
|------------------|------------|---------------|
| $C(x)$           | Generic    | 18.3%         |
| $C(x)$           | Structured | 11.8%         |
| $[C(x) - \alpha]_+ - \delta$ | Generic | 7.6%         |
| $[C(x) - \alpha]_+ - \delta$ | Structured | 3.7%         |

Table 2: Accuracy of CVaR upper bound.