DCE-MRI of locally-advanced carcinoma of the uterine cervix: Tofts analysis versus non-model-based analyses

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Abstract

Background: Dynamic contrast-enhanced magnetic resonance imaging (DCE-MRI) may provide biomarkers of the outcome of locally-advanced cervical carcinoma (LACC). There is, however, no agreement on how DCE-MRI recordings should be analyzed. Previously, we have analyzed DCE-MRI data of LACC using non-model-based strategies. In the current study, we analyzed DCE-MRI data of LACC using the Tofts pharmacokinetic model, and the biomarkers derived from this analysis were compared with those derived from the non-model-based analyses.

Methods: Eighty LACC patients given cisplatin-based chemoradiotherapy with curative intent were included in the study. Treatment outcome was recorded as disease-free survival (DFS) and overall survival (OS). DCE-MRI series were analyzed voxelwise to produce $K^{\text{trans}}$ and $v_e$ frequency distributions, and ROC analysis was used to identify the parameters of the frequency distributions having the greatest potential as biomarkers. The prognostic power of these parameters was compared with that of the non-model-based parameters LETV (low-enhancing tumor volume) and TVIS (tumor volume with increasing signal).

Results: Poor DFS and OS were associated with low values of $K^{\text{trans}}$, whereas there was no association between treatment outcome and $v_e$. The $K^{\text{trans}}$ parameters having the greatest prognostic value were p35-$K^{\text{trans}}$ (the $K^{\text{trans}}$ value at the 35 percentile of a frequency distribution) and RV-$K^{\text{trans}}$ (the tumor subvolume with $K^{\text{trans}}$ values below 0.13 min$^{-1}$). Multivariate analysis including clinical parameters and p35-$K^{\text{trans}}$ or RV-$K^{\text{trans}}$ revealed that RV-$K^{\text{trans}}$ was the only independent prognostic factor of DFS and OS. There were significant correlations between RV-$K^{\text{trans}}$ and LETV and between RV-$K^{\text{trans}}$ and TVIS, and the prognostic power of RV-$K^{\text{trans}}$ was similar to that of LETV and TVIS.

Conclusions: Biomarkers of the outcome of LACC can be provided by analyzing DCE-MRI series using the Tofts pharmacokinetic model. However, these biomarkers do not appear to have greater prognostic value than biomarkers determined by non-model-based analyses.

Keywords: Cervical carcinoma, Biomarkers, DCE-MRI, Tofts pharmacokinetic model
Background
Tumor hypoxia is a major cause of treatment failure in patients with locally-advanced cervical carcinoma (LACC) given cisplatin-based chemoradiotherapy [1–3]. The outcome of LACC may be improved by personalizing the treatment, and personalized therapy of LACC requires novel biomarkers of treatment outcome [4]. Dynamic contrast-enhanced magnetic resonance imaging (DCE-MRI) may provide information on the extent of hypoxia in LACC [4–7], and the outcome of LACC has been shown to be associated with parameters derived from DCE-MRI data [8–12]. However, these associations are fairly weak, and a sound DCE-MRI strategy for predicting the outcome of LACC has yet to be developed.

We have analyzed DCE-MRI data of LACC voxelwise and identified two biomarkers of disease-free survival (DFS) and overall survival (OS): low-enhancing tumor volume (LETV) [13] and tumor volume with increasing signal intensity (TVIS) [14]. LETV refers to the tumor volume showing low contrast enhancement during the first 60 s of the recordings, and TVIS represents the tumor volume showing increasing contrast enhancement during a 6-min-long interval in the late-phase of the DCE-MRI series [13, 14]. The calculation of LETV and TVIS was not based on any pharmacokinetic model.

Two pharmacokinetic models are used frequently to analyze clinical DCE-MRI series: the Brix model [15, 16] and the Tofts model [17, 18]. These models are based on physiological properties of tumors, and consequently, it may be preferable to analyze DCE-MR recordings of tumors using model-based strategies rather than non-model-based strategies [19]. In an earlier study of LACC, we compared the prognostic power of the parameters of the Brix model with that of LETV and TVIS and concluded that biomarkers derived from Brix analysis are not superior to LETV and TVIS [20].

In the present investigation, the DCE-MR recordings of the LACC patients included in our studies of LETV and TVIS were analyzed using the Tofts model. The Tofts model has the advantage to the Brix model that the analysis can be based on the concentration of contrast agent in the tumor tissue rather than signal intensity [15–18]. The objective of the investigation was to determine whether the parameters of the Tofts model may be clinically useful biomarkers of the outcome of LACC and to compare the prognostic value of these biomarkers with that of LETV and TVIS.

Materials and methods
Patients
Eighty patients with untreated LACC (FIGO stage IB through IVA) admitted to the Norwegian Radium Hospital were included in the study. The demographics of the patients have been described earlier [13, 14]. After DCE-MRI, the patients were treated with concurrent cisplatin-based chemoradiotherapy with curative intent. Briefly, external beam radiation therapy was given in 25 fractions during a period of 5 weeks to a total dose of 50 Gy to the primary tumor, parametria, and adjacent pelvic wall and 45 Gy to the rest of the pelvic region. In addition, 5–6 fractions of intracavitary brachytherapy with a dose of 4.2 Gy per fraction were given to Point A. Chemotherapy with cisplatin (40 mg/m²) was given weekly with a maximum of 6 courses during the radiation therapy period.

The primary endpoints were DFS, defined as the time from diagnosis to local or distant relapse or death, and OS, defined as the time from diagnosis to death. The study was approved by the regional committee of medical research ethics in southern Norway and was carried out in accordance with the Declaration of Helsinki. Informed consent was obtained from all patients.

MRI
The MRI protocol has been described in detail elsewhere [13]. Briefly, the pelvis was scanned with axial $T_2$-weighted fast spin echo sequences, and tumor volume and lymph node status were assessed by examining the $T_2$-weighted images in the dicom viewer Osirix. A gradient echo sequence was used to generate axial proton density-weighted images prior to and after DCE-MRI. DCE-MRI was conducted by using an axial $T_1$-weighted spoiled gradient recalled sequence [repetition time: 160 ms; echo time: 3.5 ms; flip angle: 90°; number of excitations: 1; field of view: 20 × 20 cm²; image matrix: 256 × 256; slice thickness: 5 mm; slice spacing: 6 mm; voxel size: 0.78 × 0.78 × 5.0 mm³; temporal resolution: 29 s; sampling time: 10 min]. Three precontrast images were acquired, and contrast agent was injected manually in a bolus dose of 0.1 mmol/kg Gd-DTPA during a period of ~3 s. To enable calculation of the concentration of Gd-DTPA in the tumor tissue, a calibration tube with two chambers, one filled with saline and the other with a 0.5-mmol Gd-DTPA solution, was inserted into the vagina prior to the imaging. The tube was positioned as close to the tumor as possible, and acceptable positioning was verified by MRI.

Image analysis
Voxel-by-voxel analysis of DCE-MR images was conducted by using software developed in Matlab. Minor movements of the tumor tissue during image acquisition were corrected for by coordinate mapping. To avoid significant influence of body cavities and tumor necrosis, voxels showing signal intensities in $T_2$-weighted images consistent with the presence of air or water were excluded from analysis. Gd-DTPA concentrations were
calculated from signal intensities using the method of Hittmair et al. [21], as described in detail elsewhere [22].

Plots of Gd-DTPA concentration versus time after contrast injection were generated, and the Levenberg-Marquardt least squares minimization method was used to fit curves to the data, using the standard Tofts equation [18]:

\[ C_t(T) = K^{\text{trans}} \int_0^T C_p(t) e^{-\frac{K^{\text{trans}}(t-t)}{v_c}} \, dt. \]

Here, \( C_t(T) \) is the tumor tissue concentration of Gd-DTPA at time \( T \), \( K^{\text{trans}} \) is the volume transfer constant of Gd-DTPA, \( v_c \) is the fractional distribution volume of Gd-DTPA in the tumor tissue, and \( C_p(t) \) is the arterial input function:

\[ C_p(t) = A e^{-Bt} + C e^{-Dt}, \]

where \( C_p(t) \) is the plasma concentration of Gd-DTPA at time \( t \), \( A = 5.10 \, \text{mM} \), \( B = 14.2 \, \text{s}^{-1} \), \( C = 0.99 \, \text{mM} \), and \( D = 0.159 \, \text{s}^{-1} \). Frequency distributions and parametric images of \( K^{\text{trans}} \) and \( v_c \) were generated by using Sigma-Plot software.

**Statistical analysis**

Because standard first-line treatment fails in 60–70% of the patients diagnosed with LACC [23], associations between Tofts parameters and DFS or OS were studied by dividing the patient cohort into two groups consisting of one-third and two-thirds of the patients [13]. Thus, the DFS and OS of the 26 patients with the lowest \( K^{\text{trans}} \) values were compared with the DFS and OS of the 54 patients with the highest \( K^{\text{trans}} \) values, and similar comparisons were carried out for \( v_c \). Kaplan–Meier curves were compared by using the log-rank test. Univariate and multivariate Cox proportional hazard analyses were used to evaluate the prognostic power of clinical and DCE-MRI–derived parameters. The Spearman rank order correlation test was used to search for correlations between parameters. Probability values of \( p < 0.05 \) were considered significant.

**Results**

Relevant anatomical MR images of a representative LACC patient are presented in Fig. 1. The position and signal intensities of the calibration tube are illustrated in a sagittal and two axial precontrast \( T_1 \)-weighted images (Fig. 1a), and a proton density-weighted image, a precontrast \( T_1 \)-weighted image, and a postcontrast \( T_1 \)-weighted image are shown to illustrate the signal intensities of the tumor tissue (Fig. 1b).

Tumor signal enhancement after Gd-DTPA administration differed substantially among individual patients, and Fig. 2 presents \( K^{\text{trans}} \) data for two representative patients, one with a high-enhancing tumor (Fig. 2a) and the other with a low-enhancing tumor (Fig. 2b). The Tofts model was seen to be well suited for analyzing the DCE-MRI series. Single-voxel plots of Gd-DTPA concentration versus time were well fitted by the model, and the numeric values of \( K^{\text{trans}} \) were generally higher in high-enhancing than in low-enhancing tumors. However, the intratumor heterogeneity in \( K^{\text{trans}} \) was substantial, as illustrated by the \( K^{\text{trans}} \) frequency distributions and parametric maps in Fig. 2.

Detailed analyses of the \( K^{\text{trans}} \) frequency distributions were carried out to investigate whether \( K^{\text{trans}} \) may provide valid prognostic information on DFS and OS. First, the \( K^{\text{trans}} \) values at each integer percentile of the frequency distributions were determined, and a log-rank test was conducted for each percentile to examine whether the \( K^{\text{trans}} \) values at any of the percentiles could discriminate between the two patient groups. Plots of the log-rank \( p \) value versus \( K^{\text{trans}} \) percentile were established, and the plots revealed that the \( K^{\text{trans}} \) values at percentiles above 30 provided \( p \) values < 0.05 for both DFS and OS (Fig. 3a). Receiver operating characteristic (ROC) analysis was carried out to determine the optimal \( K^{\text{trans}} \) percentile. The area under the ROC-curve is plotted versus \( K^{\text{trans}} \) percentile in Fig. 3b, and the optimal percentile, corresponding the largest area under the ROC-curve, was found to be the 35 percentile for DFS as well as OS. The \( K^{\text{trans}} \) value at this percentile was termed 35p-\( K^{\text{trans}} \).

Next, we calculated the number of voxels and the corresponding tumor volumes with \( K^{\text{trans}} \) values below a wide range of threshold values up to 0.50 \text{min}^{-1}, using increments of 0.0025 \text{min}^{-1}. A log-rank test was carried out for each threshold value to investigate whether the tumor volumes at any of the threshold values could discriminate between the two patient groups. Plots of the log-rank \( p \) value versus \( K^{\text{trans}} \) threshold value were generated for DFS and OS, and these plots showed that \( K^{\text{trans}} \) threshold values between 0.07 \text{min}^{-1} and 0.24 \text{min}^{-1} provided \( p \) values < 0.05 for both endpoints (Fig. 3c). ROC analysis revealed that the optimal \( K^{\text{trans}} \) threshold value was 0.13 \text{min}^{-1}, regardless of whether DFS or OS was considered (Fig. 3d). The tumor volume with \( K^{\text{trans}} \) values below 0.13 \text{min}^{-1} was presumed to represent the risk volume (RV) and was termed RV-\( K^{\text{trans}} \).

Binary tumor maps showing the distribution of voxels with \( K^{\text{trans}} \) values above or below 0.13 \text{min}^{-1} are included in Fig. 2. In general, the voxels constituting RV-\( K^{\text{trans}} \) were contiguous in tumors having a large RV-\( K^{\text{trans}} \) (Fig. 2b).

Kaplan–Meier plots for DFS and OS based on 35p-\( K^{\text{trans}} \) and RV-\( K^{\text{trans}} \) are presented in Fig. 4. The patients with low \( K^{\text{trans}} \) values did worse than those with high \( K^{\text{trans}} \) values. The 5-year survival rates of the patients
with low 35p-$K_{trans}$ values and those with high 35p-$K_{trans}$ values were 44 and 72%, respectively (DFS: $p = 0.032$) and 46 and 76%, respectively (OS: $p = 0.0065$). Similarly, the patients with large RV-$K_{trans}$ and those with small RV-$K_{trans}$ showed 5-year DFS rates of 41 and 73%, respectively ($p = 0.0044$) and 5-year OS rates of 42 and 78%, respectively ($p = 0.0034$).

Univariate Cox regression analysis showed that 35p-$K_{trans}$ and RV-$K_{trans}$ had significant impact on OS, and that RV-$K_{trans}$ but not 35p-$K_{trans}$ had significant impact on DFS (Table 1). A similar analysis of clinical parameters revealed that DFS and OS were influenced significantly by tumor volume and FIGO stage, but not by lymph node status, tumor histology, and patient age (Table 1). Furthermore, RV-$K_{trans}$ was found to be the only independent prognostic factor of DFS and OS in multivariate Cox regression analysis involving tumor volume, FIGO stage, lymph node status, and 35p-$K_{trans}$ or RV-$K_{trans}$ (Table 1).

The DCE-MR recordings of the patients included in this study have been subjected to non-model-based analyses previously, and LETV and TVIS were identified as independent prognostic factors [13, 14]. LETV and TVIS represent tumor RVs, and plots of RV-$K_{trans}$ versus LETV and TVIS revealed strong correlations between the RV derived from the Tofts analysis and the RVs derived from the non-model-based analyses ($p < 0.0001$; Fig. 5). The horizontal and vertical lines in Fig. 5 represent the border between small and large RVs, and show that the majority of the patients were stratified into the same risk group by the model-based and non-model-based RVs.

By subjecting the $v_e$ frequency distributions of the tumors to analyses similar to those described above for the $K_{trans}$ frequency distributions, it was seen that $v_e$ did not provide valid prognostic information on DFS or OS. Parametric $v_e$ images and $v_e$ frequency distributions of the tumors used as examples in Fig. 2, and plots of log-rank $p$ value versus $v_e$ percentile and log-rank $p$ value...
Fig. 3 Log-rank and ROC analyses. The values of $K^{\text{trans}}$ at each percentile of the frequency distributions and the tumor volumes (RV-$K^{\text{trans}}$) with $K^{\text{trans}}$ values below a wide range of threshold values were calculated for the 80 tumors included in the study, and for each $K^{\text{trans}}$ percentile and each $K^{\text{trans}}$ threshold value, the outcome of the patients with high values of $K^{\text{trans}}$ or RV-$K^{\text{trans}}$ was compared with the outcome of those with low values, using the log-rank test with DFS and OS as endpoints. ROC analysis was carried out to identify the $K^{\text{trans}}$ percentile and $K^{\text{trans}}$ threshold value with the highest discriminative power. 

a Log-rank $p$ value versus $K^{\text{trans}}$ percentile. 

b Area under ROC-curve versus $K^{\text{trans}}$ percentile. 

c Log-rank $p$ value versus $K^{\text{trans}}$ threshold value. 

d Area under ROC-curve versus $K^{\text{trans}}$ threshold value.
Fig. 4 Treatment outcome. Kaplan–Meier curves for DFS and OS of LACC patients stratified by 35p-K\text{trans} a,b and RV-K\text{trans} c,d. p values: log-rank test

Table 1 Cox regression analysis of clinical and DCE-MRI-derived parameters

|                | Disease-free survival | Overall survival | Disease-free survival | Overall survival |
|----------------|-----------------------|------------------|-----------------------|------------------|
|                | p value               | p value          | p values\textsuperscript{a} | p values\textsuperscript{a} |
| Tumor volume   | 0.026                 | 0.019            | 0.49                  | 0.69             | 0.18                  | 0.76 |
| FIGO stage     | 0.0057                | 0.017            | 0.065                 | 0.055            | 0.099                 | 0.091 |
| Lymph node status | 0.13               | 0.36             | 0.73                  | 0.84             | 0.81                  | 0.60 |
| Tumor histology | 0.30                 | 0.31             | –                     | –                | –                     | –    |
| Patient age    | 0.30                  | 0.17             | –                     | –                | –                     | –    |
| 35p-K\text{trans} | 0.15               | 0.047            | 0.20                  | –                | 0.057                 | –    |
| RV-K\text{trans} | 0.00033            | < 0.00001        | –                     | 0.026            | –                     | 0.0015 |

Values of p < 0.05 are highlighted in bold
\textsuperscript{a}p values refer to multivariate regression analyses including tumor volume, FIGO stage, lymph node status, and either 35p-K\text{trans} or RV-K\text{trans}

35p-K\text{trans}, the value of K\text{trans} at the 35th percentile of the K\text{trans} frequency distribution of a tumor
RV-K\text{trans}, the tumor subvolume with K\text{trans} values < 0.13 min\textsuperscript{-1}
versus $v_e$ threshold value are presented in Additional file 1: Figure S1.

**Discussion**

Imaging plays an important role in the diagnostics, radiation therapy planning, and treatment monitoring of LACC [4]. In addition to MRI, positron emission tomography (PET)/computed tomography (CT) has shown promise for hypoxia imaging. Current studies are investigating the potential of several PET hypoxia tracers, including $^{18}$F-fluoromisonidazole, $^{18}$F-fluoroerythronitroimidazole, $^{18}$F-fluoroazomycin-arabinoside, and $^{60}$Cu-diacetyl-bis(N-methylthiosermicarbazon), and interesting observations have been reported [24]. It has also been hypothesized that $^{18}$F-fluorodeoxyglucose ($^{18}$F-FDG) may be a clinically useful surrogate hypoxia tracer [24]. If this hypothesis proves to be valid, $^{18}$F-FDG PET/CT may turn out to be a powerful imaging modality for LACC, since $^{18}$F-FDG PET/CT already is used successfully to detect pathological lymph nodes and high-risk radiation therapy targets [25, 26].

The potential of DCE-MRI as an imaging modality for providing hypoxia-associated biomarkers of the outcome of LACC was investigated in the study reported in this communication. DCE-MRI series of LACC were analyzed by using the Tofts pharmacokinetic model, and the prognostic power of biomarkers derived from the Tofts analysis was compared with that of biomarkers derived from non-model-based analyses. Tofts analysis has the advantage to non-model-based analyses that it is based on the physiology of the imaged tumor, but to utilize this advantage in full, information on the concentration of contrast agent in the tissue is required. This information was provided in the present study by using a two-chamber vaginal calibration tube.

Significant associations were found between $p35-K_{\text{trans}}$ or RV-$K_{\text{trans}}$ on the one hand side and DFS or OS on the other, as revealed by univariate Cox regression analysis. Multivariate analysis including clinical parameters and $p35-K_{\text{trans}}$ or RV-$K_{\text{trans}}$ showed that RV-$K_{\text{trans}}$ was a strong independent prognostic factor of DFS and OS, whereas $p35-K_{\text{trans}}$, tumor volume, FIGO stage, and lymph node status were not independent prognostic factors. These observations suggest that Tofts analysis of DCE-MRI data can provide biomarkers of the outcome of LACC. Moreover, the prognostic power of RV-$K_{\text{trans}}$ is strong compared with that of the $K_{\text{trans}}$ value at any percentile of the $K_{\text{trans}}$ frequency distribution.

Poor treatment outcome was associated with low values of $K_{\text{trans}}$ and, consequently, large values of RV-$K_{\text{trans}}$. Clinical investigations have revealed that low $K_{\text{trans}}$ values reflect poor vascularization and blood supply in LACC [9, 27, 28]. Preclinical studies of human cervical carcinoma xenografts have shown strong inverse correlations between median $K_{\text{trans}}$ and fraction of hypoxic tumor tissue, regardless of whether radiobiological or pimonidazole-based immunohistochemical assays were used to quantify tumor hypoxia [5, 29–31]. Therefore, it is reasonable to assume that RV-$K_{\text{trans}}$ mirrors the hypoxic subvolume in tumors of the uterine cervix. This assumption is consistent with investigations having shown that poor outcome of LACC is associated with low oxygen tension in the primary tumor, as measured with polarographic Eppendorf $pO_2$ electrodes [1–3].

RV-$K_{\text{trans}}$ correlated strongly with the non-model-based DCE-MRI-derived parameters LETV and TVIS, and the vast majority of the patients were stratified into
the same risk group by RV-$K^{\text{trans}}$, LETV, and TVIS. Furthermore, the RV-$K^{\text{trans}}$-based Kaplan-Meier plots for DFS and OS reported here were similar to the LETV-based and TVIS-based Kaplan-Meier plots reported previously [13, 14]. Consequently, the prognostic power of biomarkers identified by Tofts analysis does not appear to be stronger than that of biomarkers identified by non-model-based analyses.

RV-$K^{\text{trans}}$, LETV, and TVIS have specific disadvantages and advantages as biomarkers of LACC. RV-$K^{\text{trans}}$ has the disadvantage that an arterial input function and long scanning times are needed for its assessment. The advantage of RV-$K^{\text{trans}}$ is that it is related to the extent of tumor hypoxia. The assessment of LETV requires signal intensity-dependent threshold values [13], and because signal intensities vary depending upon MR protocol and scanner, radiology departments have to establish their own threshold values. The advantage of LETV is that short scanning times of only 1–2 min are needed, and this advantage is of significant importance because scanning time is a major limiting factor in many centers. Similar to the calculation of RV-$K^{\text{trans}}$, the calculation of TVIS requires long scanning times, but the advantage of TVIS is that the calculation is independent of signal intensity, thus facilitating comparisons of results across institutions.

The current investigation has some limitations. First, $T_2$-weighted images were used to delineate tumors and to exclude image regions influenced by body cavities and tumor necrosis, and despite the fact that these tasks were carried out by two experienced radiologists, there may have been some uncertainties in the assessment of the regions of interest. Second, the MR images were recorded several years ago when the spatial and temporal resolutions of the recordings were poorer than today’s standard due to scanner limitations. Third, the Tofts analysis was performed by using a population-based arterial input function rather than individual arterial input functions. Fourth, only 80 patients were included in the study, and consequently, it requires validation in an independent patient cohort. Despite these limitations, important conclusions can be drawn from our investigation.

**Conclusions**

By subjecting DCE-MRI series of LACC to pharmacokinetic analysis using the Tofts model, prognostic factors independent of well-established clinical prognostic factors can be provided. Furthermore, RV-$K^{\text{trans}}$ has stronger prognostic power than the $K^{\text{trans}}$ value at any percentile of the $K^{\text{trans}}$ frequency distribution. However, the prognostic value of RV-$K^{\text{trans}}$ is not necessarily superior to the prognostic value of RVs assessed by non-model-based analyses.

**Supplementary information**

Supplementary information accompanies this paper at https://doi.org/10.1186/s13014-020-01526-2.

**Additional file 1.** DCE-MRI of locally-advanced carcinoma of the uterine cervix: Tofts analysis versus non-model-based analyses

**Abbreviations**

DCE-MRI: Dynamic contrast-enhanced magnetic resonance imaging; DFS: Disease-free survival; Gd-DTPA: Gadolinium diethylenetriamine pentaacetic acid; LACC: Locally-advanced cervical carcinoma; LETV: Low-enhancing tumor volume; OS: Overall survival; ROC: Receiver operating characteristic; RV: Risk volume; TVIS: Tumor volume with increasing signal intensity

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**Authors’ contributions**

KVL and TGS analyzed DCE-MRI data, performed statistical analysis, and drafted the manuscript. GBK collected and analyzed clinical outcome data. EKR designed the study, supervised data analysis, and prepared the final manuscript. The authors read and approved the final manuscript.

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**Availability of data and materials**

The datasets used and/or analyzed during the current study are available from the corresponding author on reasonable requests.

**Ethics approval and consent to participate**

The study was approved by the regional committee of medical research ethics in southern Norway and was carried out in accordance with the Declaration of Helsinki. Informed consent was obtained from all patients.

**Consent for publication**

Not applicable.

**Competing interests**

The authors declare that they have no competing interests.

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