Cholesterol testing among men and women with disability: the role of morbidity

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Purpose: Despite more frequent use of health services by people living with disability, the quality of preventive care received may be suboptimal. In this retrospective cohort study, we used administrative data to examine the relationship between cholesterol testing and levels of disability and morbidity among women and men in Ontario, Canada.

Methods: We linked multiple provincial-level databases in this study. In stratified analyses for women and men, we used multivariable logistic regression to examine differences in cholesterol testing, and we tested for an interaction effect between disability and morbidity. In a secondary analysis, we tested for a three-way interaction between sex, disability, and morbidity on the entire cohort.

Results: There was an interaction between morbidity and disability for both women and men. Women and men with no chronic conditions appeared to be least likely to be up-to-date on cholesterol testing, and among this group, those with moderate disability were more likely to be up-to-date on cholesterol testing than those with no disability (adjusted odds ratio [AOR] =1.51; 95% confidence interval [CI] 1.20–1.90 for women; AOR =1.16; 95% CI 1.00–1.34 for men). Among women and men who had one chronic condition, having severe disability put them at significant disadvantage versus those with no disability. Only 58.5% of men with no disability and no chronic conditions were up-to-date on cholesterol testing.

Conclusion: An intermediate level of health care need (reflected in this study as level of disability and level of morbidity) may provide a benefit for cholesterol testing, and conversely, health care needs that are too few or too great may negatively affect testing. Public health and practice-based interventions need to be explored to address these findings.

Keywords: primary care, preventative care, cholesterol testing, disability, morbidity

Introduction
Disability has been defined as a limitation to performing everyday activities because of a health condition.1,2 Despite more frequent use of health services by people living with disability, the quality of care that these individuals receive may be inferior to that of the general population, particularly with regards to preventive care.3–10

The relationship between the receipt of preventive care and disability may be complex, with multi-morbidity (ie, the coexistence of at least two chronic conditions) being to play a role. People with disability tend to have higher levels of comorbidity, and morbidity has been associated with such measures of preventive care as unplanned hospital admissions, breast cancer screening, and cervical cancer screening.9,10,12–14 In previous work, we found an interaction between the levels of disability and morbidity...
for cervical cancer screening, whereby women with a higher level of disability were less likely to be screened than women with a lower level of disability as their level of morbidity increased.

For breast cancer screening, we observed an inverse V-shaped relationship between screening and levels of disability. Women of moderate disability had higher levels of screening than women with no disability (adjusted odds ratio [AOR] = 1.20; 95% confidence interval [CI] 1.09–1.38) and women with severe disability had lower odds of being screened compared to women with moderate disability (AOR 0.72; 95% CI 0.63–0.82). Women with one chronic condition had higher odds of screening than women with no chronic conditions (AOR 1.31; 95% CI 1.17–1.46).10

However, both breast and cervical cancer screening involve procedure-based testing and are limited to women. We know of no studies that have examined the relationship between disability, morbidity, and simpler common secondary prevention tests, such as cholesterol testing performed by venipuncture, for both men and women in Ontario. Cholesterol testing is commonly performed in the primary-care setting and has clear guidelines on testing.15,16 Therefore, in this retrospective cohort study, we used provincial-level administrative data to conduct a stratified analysis by sex, examining the relationship between adherence to guidelines for cholesterol testing and levels of disability and morbidity among screen-eligible women and men in Ontario, Canada.

Methods

Data sources

We accessed multiple provincial data sources at the Institute for Clinical Evaluative Sciences (ICES), for this study. The 2005 and 2007/2008 Canadian Community Health Surveys (CCHS) were used to identify the cohort. The CCHS is a cross-sectional self-report survey administered by Statistics Canada that gathers information on sociodemographics, health status, and health determinants for the Canadian population. Using a unique encoded identifying number that allows individuals to be tracked through multiple databases, we linked respondents from the two aforementioned cycles of CCHS with several administrative health databases. The Ontario Health Insurance Plan (OHIP) Claims contains physicians’ fee-for-service claims for Ontario’s single-payer health insurance plan, and the Registered Persons Database contains basic demographic and vital statistics information for each Ontario resident.17 The Canadian Institute of Health Information Discharge Abstract Database contains fee codes and corresponding diagnostic codes claimed by Ontario’s physicians for in-hospital stays or procedures. The Ontario Myocardial Infarction Database (OMID) identifies all patients with an incident hospital admission for acute myocardial infarction. These datasets were linked using unique encoded identifiers and analyzed at ICES. This study was approved by the institutional review board at Sunnybrook Health Sciences Centre, Toronto, ON, Canada. Patient consent was not required as this was a retrospective study and it would have been infeasible to obtain consent. No direct identifiers were available to the research team.

Study population

The study population was drawn from respondents to the 2005 CCHS and 2007/2008 CCHS who agreed to have their responses linked with their personal health information at ICES (~30,000 people per CCHS cycle). In order to be included in the study, people needed to be residents of Ontario, alive during the entire observation window, eligible for OHIP during the entire observation window, and had to have answered the Restriction of Activities: Participation and Activity Limitation (RACDP AL) questions on the CCHS. The observation window was defined as the 3 calendar years after completion of the CCHS (ie, January 1, 2006–December 31, 2008, and January 1, 2009–December 31, 2011). Male cohort members needed to be 40–74 years of age and females needed to be 50–74 years of age, during the observation window. These age limits were based on Canadian Cardiovascular Society lipid guidelines.15 We excluded individuals who had a diagnosis of myocardial infarction, as per the OMID, at any time on or before the last day of the observation window, as these individuals would not be eligible for cholesterol testing as a preventative measure for the entire 3-year observation window.18

Study measures

We defined cholesterol testing by a binary outcome of whether individuals were or were not up-to-date on cholesterol testing as per guidelines,15 ie, having had at least one cholesterol test conducted during the 3-year observation window. The Canadian Cardiovascular Society guidelines recommend cholesterol testing for men who are 40+ years of age and women who are 50+ years of age, and repeating every 3–5 years for those with a Framingham risk score of less than 5% and annually for those with a score of 5% or greater.11 As we were unable to know participants’ Framingham scores with available data, we chose a 3-year time frame. These guidelines are the most commonly used by Canadian physicians.16 We used OHIP fee codes to identify cholesterol testing.

We used the RACDP AL items from the CCHS to define disability. The RACDP AL is based on the response to five questions and classifies respondents by the frequency with which the
they experience activity limitations due to a long-term physical and/or mental health problem that has lasted, or is expected to last 6 months or more. The activity limitations relate to the following questions: 1) Do you have difficulty hearing, seeing, communicating, walking, climbing stairs, bending, learning or doing any similar activities? 2) Does a long-term physical or mental condition or health problem reduce the amount or the kind of activity you can do at home, at work, and at school? 3) Does a long-term physical or mental condition or health problem reduce the amount or the kind of activity you can do in other activities, for example, transportation or leisure? Response options are “often”, “sometimes”, and “never”. If participants answered “never” to all five RACDPAL items, we classified them as having “no disability”, and they were used as the reference group. Participants were defined as having “severe disability” if they answered “often” to at least one question. All other participants were classified as having “moderate disability”9,10. The RACDPAL items were answered by more than 99% of respondents on both CCHS cycles used.

We defined the level of morbidity on the basis of the self-reported presence of at least one of several chronic conditions noted on the CCHS, namely arthritis, hypertension, COPD, diabetes, heart disease, cancer, stroke, Alzheimer’s or other dementias, and mood/anxiety disorder. More than 40% of Canadian adults have reported having at least one of these conditions.19 We categorized participants as having zero, one, or at least two of these chronic conditions, where participants with zero chronic conditions were used as the reference group.

We also included sociodemographic measures from the CCHS, specifically age, ethnicity (country of birth), level of education, and household income. The Rurality Index of Ontario score based on postal codes of residence and the number of physician visits (both primary care and specialist) during the study period was further obtained from administrative databases.20

Data analysis
In stratified analyses for women and men, we conducted descriptive statistics to describe demographics of the study cohort, as well as univariate and bivariate analyses, where all statistical tests were performed at the 5% level of significance, two-sided, using SAS for Unix, version 9.1.3 (SAS Institute, Cary, NC). We used multivariable logistic regression to examine differences in cholesterol testing. Variables included age as a continuous variable, household income, education, country of birth, rurality, marital status, level of disability, and level of morbidity, and we tested for an interaction effect between disability and morbidity. We excluded participants with missing data from the bivariate and multivariable analyses. All analyses were stratified by sex. In a secondary analysis, we also tested for a three-way interaction between sex, level of morbidity, and level of disability on the entire study cohort, to assess if the effect was different for women versus men.

Results
Overall, 67.1% of men (n=9,313) were tested in their 3-year observation window as compared to 77.0% of women (n=8,725). Table 1 describes the demographic characteristics of the 11,335 women and 13,883 men included in the study cohort. Among both women and men, those with disability had lower income, higher morbidity, and more physician contact during the study period. Women appeared to be more likely to be of lower income and of higher morbidity and to have more physician contact than men, among those both with and without disability. Almost half (48.2%) of women with disability had two or more chronic conditions as compared to 33.0% of men with disability.

Table 2 shows the results of the bivariate analyses. Across sociodemographic categories, women and men with moderate disability tended to have the highest risk of up-to-date testing and women and men with no disability tended to have the lowest. A notable exception was among those with one chronic condition and with two or more chronic conditions, where no disability was associated with the highest cholesterol testing and severe disability with the lowest.

When looking at the cross-tabulations of level of morbidity and level of disability (Table 2), for both women and men, the highest risk of up-to-date testing was for those with two or more chronic conditions but no disability (83.4% and 83.3%, respectively), and the lowest risk was for those with no chronic conditions and no disability (69.2% and 58.5%, respectively). For both sexes, level of morbidity seemed to act as an effect modifier on the relationship between level of disability and cholesterol testing. Although the risk of up-to-date testing decreased as level of disability increased for those with at least one chronic condition, this did not hold true for those with no chronic conditions. For those women and men with no chronic conditions, up-to-date testing was highest among those with moderate disability.

In multivariable analyses (Figure 1), the interaction between disability and morbidity remained after adjusting for other variables. For both sexes, a moderate level of disability (versus no disability) was associated with higher adjusted odds of cholesterol testing for those with no chronic conditions. However, this did not hold true for those with some
Table 1 | Demographic characteristics of 11,335 women and 13,883 men in the study cohort

|                | No disability | Moderate disability | Severe disability |
|----------------|---------------|---------------------|-------------------|
| **Women**      |               |                     |                   |
| **Age**        | N=6,225       | N=2,580             | N=2,530           |
| Mean ± SD median (IQR) | 59.7±6.1 | 60.4±6.1            | 60.3±6.1          |
| Country of birth | Canada | 4,641 (74.4%)      | 1,985 (76.9%)     | 2,046 (80.9%)     |
| Other          | 1,455 (23.4%) | 541 (21.0%)        | 451 (17.8%)       |
| Missing        | 129 (2.1%)   | 54 (2.1%)           | 33 (1.3%)         |
| Marital status | Married/common-law | 4,138 (66.5%) | 1,636 (63.4%) | 1,406 (55.6%) |
| Widowed/single | 1,181 (19.0%) | 513 (19.9%)        | 568 (22.5%)       |
| Divorced/separated | 903 (14.5%) | 430 (16.7%)        | 554 (21.9%)       |
| Missing        | <5 (n/a)     | <5 (n/a)            | <5 (n/a)          |
| Education      | <Secondary | 1,085 (17.4%)      | 529 (20.5%)       | 627 (24.8%)       |
| Secondary or some post-secondary | 1,769 (28.4%) | 643 (24.9%)   | 616 (24.3%)       |
| Post-secondary | 3,350 (53.8%) | 1,399 (54.2%)     | 1,277 (50.5%)     |
| Missing        | 21 (0.3%)    | 9 (0.3%)            | 10 (0.4%)         |
| Household income | <$30,000 | 1,112 (17.9%)      | 607 (23.5%)       | 907 (35.8%)       |
| $30,000 to <$59,999 | 1,935 (31.1%) | 806 (31.2%)        | 766 (30.3%)       |
| $60,000 to <$99,999 | 1,436 (23.1%) | 573 (22.2%)       | 404 (16.0%)       |
| $100,000+      | 1,106 (17.8%) | 360 (14.0%)        | 231 (9.1%)        |
| Missing        | 636 (10.2%)  | 234 (9.1%)          | 222 (8.8%)        |
| Rurality Index of Ontario score | 0–9 (large urban) | 3,006 (48.3%) | 1,228 (47.6%)   | 1,171 (46.3%)     |
| 10–44 (small urban) | 2,261 (36.3%) | 966 (37.4%)       | 988 (39.1%)       |
| 45+ (rural)    | 867 (13.9%)  | 343 (13.3%)         | 345 (13.6%)       |
| Missing        | 91 (1.5%)    | 43 (1.7%)           | 26 (1.0%)         |
| Morbidity      | No chronic conditions | 2,913 (46.8%) | 563 (21.8%) | 345 (13.6%) |
| 1              | 1,118 (18.0%) | 977 (37.9%)        | 763 (30.2%)       |
| 2+             | 2,194 (35.2%) | 1,040 (40.3%)      | 1,422 (56.2%)     |
| Missing        | <5 (n/a)     | <5 (n/a)            | <5 (n/a)          |
| Number of physician visits during study period | 345 (13.6%) | 345 (13.6%) | 345 (13.6%) |
| Median (IQR)   | 13 (7–22)    | 18 (10–28)          | 23 (13–37)        |
| **Men**        |               |                     |                   |
| **Age**        | N=8,528       | N=2,832             | N=2,523           |
| Mean ± SD median (IQR) | 53.5±9.1 | 55.1±9.0            | 56.0±8.7          |
| Country of birth | Canada | 6,370 (74.7%)      | 2,269 (80.1%)     | 2,076 (82.3%)     |
| Other          | 1,940 (22.7%) | 519 (18.3%)        | 406 (16.1%)       |
| Missing        | 218 (2.6%)   | 44 (1.6%)           | 41 (1.6%)         |
| Marital status | Married/common-law | 6,096 (71.5%) | 1,930 (68.1%)        | 1,609 (63.8%) |
| Widowed/single | 1,316 (15.4%) | 502 (17.7%)        | 434 (17.2%)       |
| Divorced/separated | 1,114 (13.1%) | 398 (14.1%)       | 475 (18.8%)       |
| Missing        | <5 (n/a)     | <5 (n/a)            | <5 (n/a)          |
| Education      | <Secondary | 1,241 (14.6%)      | 522 (18.4%)       | 590 (23.4%)       |
| Secondary or some post-secondary | 1,931 (22.6%) | 633 (22.4%) | 536 (21.2%)       |
| Post-secondary | 5,319 (62.4%) | 1,667 (58.9%)     | 1,379 (54.7%)     |
| Missing        | 37 (0.4%)    | 10 (0.4%)           | 18 (0.7%)         |
| Household income | <$30,000 | 866 (10.2%)        | 429 (15.1%)       | 678 (26.9%)       |
| $30,000 to <$59,999 | 2,133 (25.0%) | 797 (28.1%)       | 701 (27.8%)       |
| $60,000 to <$99,999 | 2,554 (29.9%) | 799 (28.2%) | 589 (23.3%)       |
| $100,000+      | 2,518 (29.5%) | 665 (23.5%)       | 405 (16.1%)       |
| Missing        | 457 (5.4%)   | 142 (5.0%)          | 150 (5.9%)        |
| Rurality Index of Ontario score | 0–9 (large urban) | 4,507 (52.8%) | 1,405 (49.6%) | 1,073 (42.5%) |
| 10–44 (small urban) | 2,929 (34.3%) | 1,035 (36.5%) | 990 (39.2%) |
| 45+ (rural)    | 989 (11.6%)  | 351 (12.4%)         | 426 (16.9%)       |
| Missing        | 102 (1.2%)   | 41 (1.4%)           | 34 (1.3%)         |
| Morbidity      | No chronic conditions | 5,311 (62.3%) | 1,099 (38.8%) | 648 (25.7%) |
| 1              | 2,265 (26.6%) | 985 (34.8%)        | 855 (33.9%)       |
| 2+             | 952 (11.2%)  | 748 (26.4%)         | 1,020 (40.4%)     |
| Missing        | <5 (n/a)     | <5 (n/a)            | <5 (n/a)          |
| Number of physician visits during study period | 12.6±13.8 | 17.6±19.2 | 23.0±23.0 |

Notes: All percentages reported are column percentages.
Abbreviations: SD, standard deviation; n/a, not applicable; IQR, interquartile range.
level of morbidity. For both sexes, severe disability (versus no disability) was associated with lower adjusted odds of cholesterol testing for those with one chronic condition.

In our secondary analysis, the three-way interaction between sex, level of disability, and level of morbidity was not significant ($P=0.97$), suggesting no difference in effect between women and men.

**Interpretation**

We have found that, in Ontario, morbidity acts as an effect modifier on the relationship between level of disability and adherence to guidelines for cholesterol testing. The number of chronic conditions appeared to differentiate those who were up-to-date most strongly, with women and men with no chronic conditions being least likely to be up-to-date on cholesterol testing. Among those with no chronic conditions, women and men living with a moderate level of disability were more likely to be up-to-date on cholesterol testing than those with no disability. Among women and men who had one chronic condition, also having a severe disability put them at a significant disadvantage compared to those with no disability. Additionally, we found that, across sexes and levels of disability and morbidity, men with no disability and no chronic conditions had the lowest risk of appropriate cholesterol testing, where only 58.5% of these men were up-to-date on testing and that Ontarians living with disability

| Table 2 Number and percentage of participants up-to-date on cholesterol testing by level of disability and sociodemographic characteristics, stratified by sex |
|---------------------------------------------------------------|
| **Women** | | | |
| Country of birth | | | |
| Canada | 3,410 (73.5%) | 1,561 (78.6%) | 1,538 (75.2%) |
| Other | 1,202 (82.6%) | 462 (85.4%) | 383 (84.9%) |
| Marital status | | | |
| Married/common-law | 3,185 (77.0%) | 1,301 (79.5%) | 1,091 (77.6%) |
| Widowed/single | 878 (74.3%) | 426 (83.0%) | 427 (75.2%) |
| Divorced/separated | 647 (71.7%) | 333 (77.4%) | 432 (78.0%) |
| Education | | | |
| <Secondary | 820 (75.6%) | 405 (76.6%) | 470 (75.0%) |
| Secondary or some post-secondary | 1,345 (76.0%) | 525 (81.6%) | 475 (77.1%) |
| Post-secondary | 2,533 (75.6%) | 1,123 (80.3%) | 999 (78.2%) |
| Household income | | | |
| <$30,000 | 822 (73.9%) | 465 (76.6%) | 684 (75.4%) |
| $30,000 to <$99,999 | 1,429 (73.9%) | 658 (81.6%) | 595 (77.7%) |
| $60,000 to <$99,999 | 1,106 (77.0%) | 475 (82.9%) | 313 (77.5%) |
| $100,000+ | 868 (78.5%) | 289 (80.3%) | 179 (77.5%) |
| Rurality Index of Ontario score | | | |
| 0–9 (large urban) | 2,447 (81.4%) | 1,044 (85.0%) | 957 (81.7%) |
| 10–44 (small urban) | 1,724 (76.2%) | 784 (81.2%) | 775 (78.4%) |
| 45+ (rural) | 489 (56.4%) | 209 (60.9%) | 203 (58.8%) |
| Morbidity | | | |
| No chronic conditions | 2,016 (69.2%) | 432 (76.7%) | 254 (73.6%) |
| 1 | 1,765 (80.4%) | 771 (78.9%) | 572 (75.0%) |
| 2+ | 932 (83.4%) | 858 (82.5%) | 1,125 (79.1%) |
| **Men** | | | |
| Country of birth | | | |
| Canada | 3,975 (62.4%) | 1,562 (68.8%) | 1,394 (67.1%) |
| Other | 1,454 (74.9%) | 402 (77.5%) | 306 (75.4%) |
| Marital status | | | |
| Married/common-law | 4,214 (69.1%) | 1,428 (74.0%) | 1,159 (72.0%) |
| Widowed/single | 725 (55.1%) | 317 (63.1%) | 257 (59.2%) |
| Divorced/separated | 643 (57.7%) | 254 (63.8%) | 310 (65.3%) |
| Education | | | |
| <Secondary | 782 (63.0%) | 355 (68.0%) | 393 (66.6%) |
| Secondary or some post-secondary | 1,253 (64.9%) | 433 (68.4%) | 363 (67.7%) |
| Post-secondary | 3,522 (66.2%) | 1,204 (72.2%) | 964 (69.9%) |
| Household income | | | |
| <$30,000 | 526 (60.7%) | 292 (68.1%) | 440 (64.9%) |
| $30,000 to <$99,999 | 1,377 (64.6%) | 539 (67.6%) | 492 (70.2%) |
| $60,000 to <$99,999 | 1,668 (65.3%) | 555 (69.5%) | 406 (68.9%) |
| $100,000+ | 1,721 (68.3%) | 505 (75.9%) | 289 (71.4%) |
| Rurality Index of Ontario score | | | |
| 0–9 (large urban) | 3,150 (69.9%) | 1,062 (75.6%) | 771 (71.9%) |
| 10–44 (small urban) | 1,894 (64.7%) | 728 (70.3%) | 707 (71.4%) |
| 45+ (rural) | 492 (49.7%) | 184 (52.4%) | 234 (54.9%) |
| Morbidity | | | |
| No chronic conditions | 3,107 (58.5%) | 680 (61.9%) | 398 (61.4%) |
| 1 | 1,683 (74.3%) | 713 (72.4%) | 554 (64.8%) |
| 2+ | 793 (83.3%) | 606 (81.0%) | 779 (76.4%) |

**Notes:** All percentages reported are column percentages.

**Abbreviations:** SD, standard deviation; n/a, not applicable; IQR, interquartile range.
A

| Variable | Odds Ratio (OR) CI |
|----------|-------------------|
| Age | 1.04 (1.03–1.05) |
| Education | | |
| Less than secondary vs post secondary | 0.84 (0.73–0.96) |
| Secondary vs post-secondary | 1.00 (0.89–1.12) |
| Household income | | |
| <$30,000 vs $100,000+ | 0.69 (0.57–0.83) |
| $30,000–59,999 vs $100,000+ | 0.77 (0.65–0.90) |
| $60,000–99,999 vs $100,000+ | 0.97 (0.83–1.14) |
| Rurality index | | |
| Small urban vs large urban | 0.83 (0.74–0.93) |
| Rural vs large urban | 0.32 (0.28–0.37) |
| Country of birth | | |
| Other vs Canada | 1.35 (1.16–1.53) |
| Marital status | | |
| Widowed/single vs married/common-law | 0.89 (0.78–1.02) |
| Separated/divorced vs married/common-law | 0.88 (0.76–1.01) |
| Disability | | |
| Moderate vs no disability | 1.51 (1.20–1.90) |
| Severe vs no disability | 1.24 (0.95–1.63) |
| Chronic conditions | | |
| 1 vs 0 conditions | 1.83 (1.58–2.12) |
| 2+ vs 0 conditions | 2.33 (1.91–2.84) |
| Disability*chronic conditions | | |
| Moderate vs no disability (0 conditions) | 1.51 (1.20–1.90) |
| Severe vs no disability (0 conditions) | 1.24 (0.95–1.63) |
| Moderate vs no disability (1 condition) | 0.99 (0.81–1.22) |
| Severe vs no disability (1 condition) | 0.77 (0.62–0.95) |
| Moderate vs no disability (2+ conditions) | 1.01 (0.78–1.30) |
| Severe vs no disability (2+ conditions) | 0.81 (0.65–1.01) |

B

| Variable | Odds Ratio (OR) CI |
|----------|-------------------|
| Age | 1.05 (1.05–1.06) |
| Education | | |
| Less than secondary vs post secondary | 0.83 (0.74–0.93) |
| Secondary vs post-secondary | 1.01 (0.92–1.12) |
| Household income | | |
| <$30,000 vs $100,000+ | 0.63 (0.54–0.73) |
| $30,000–59,999 vs $100,000+ | 0.73 (0.65–0.83) |
| $60,000–99,999 vs $100,000+ | 0.86 (0.77–0.95) |
| Rurality index | | |
| Small urban vs large urban | 0.79 (0.72–0.86) |
| Rural vs large urban | 0.37 (0.32–0.41) |
| Country of birth | | |
| Other vs Canada | 1.34 (1.21–1.49) |
| Marital status | | |
| Widowed/single vs married/common-law | 0.69 (0.62–0.78) |
| Separated/divorced vs married/common-law | 0.72 (0.64–0.81) |
| Disability | | |
| Moderate vs no disability | 1.16 (1.00–1.34) |
| Severe vs no disability | 1.16 (0.97–1.39) |
| Chronic conditions | | |
| 1 vs 0 conditions | 1.86 (1.65–2.10) |
| 2+ vs 0 conditions | 2.90 (2.38–3.54) |
| Disability*chronic conditions | | |
| Moderate vs no disability (0 conditions) | 1.16 (1.00–1.34) |
| Severe vs no disability (0 conditions) | 1.16 (0.97–1.39) |
| Moderate vs no disability (1 condition) | 0.87 (0.73–1.05) |
| Severe vs no disability (1 condition) | 0.74 (0.61–0.89) |
| Moderate vs no disability (2+ conditions) | 0.88 (0.67–1.15) |
| Severe vs no disability (2+ conditions) | 0.83 (0.65–1.07) |

Figure 1 Adjusted odds ratios (OR) and 95% confidence intervals (CI) from multivariable logistic regression, where variables included in the model were age (as a continuous variable), education, household income, rurality, country of birth, marital status, level of disability, level of morbidity (ie, number of chronic conditions), and the interaction between level of disability and level of morbidity.

Notes: Data for (A) women and (B) men. **p=** denotes interaction term.
are of lower socioeconomic status and have more chronic conditions than their peers, with women with disability being particularly affected.

Our findings suggest that an intermediate level of health care need (reflected in this study as level of disability and level of morbidity) may provide a benefit for cholesterol testing. In contrast, having health care needs that are either too few or too great seemingly has a negative effect on cholesterol testing. Although the reasons for these findings cannot be determined from these data, it is feasible that this is due to reduced health system contact for patients with few health care needs and too many competing demands at visits for patients with great health care needs. Our previous research showed that increasing health care needs led to less cervical cancer screening and, similar to what we have now found with cholesterol testing, that an intermediate level of health care need was associated with more breast cancer screening. An intermediate level of disability has previously been found to be associated with the highest level of receipt of immunization.

Taken together, our body of work suggests that the chances of receiving appropriate preventative care in general are highest if a patient meets some minimum threshold for contact with their primary-care provider, if he or she does not have too many other competing health demands, and for those with other demands, if the preventative test requires minimal time on the part of the provider, such as giving a patient a laboratory requisition or completing a mammography referral. When seeing patients with high health care needs, primary-care physicians face many struggles, including difficulty simultaneously adhering to numerous guidelines, which are generally written for sole conditions, and insufficient consultation time. Future research should explore how physician reminder systems, patient recall systems, and non-physician health providers can surmount these barriers to providing quality preventive care for complex patients. Reminder/recall systems could also be useful for patients with minimal health system contact.

In line with other studies, including research from Ontario, we found that women were more likely to be up-to-date on cholesterol testing than men. As discussed earlier, these findings may reflect more regular contact with the health care system for women than men, which we also found with this study cohort. Interestingly, we also found that men who were in a married or common-law relationship (the majority of whom would be in a heterosexual relationship) were more likely to be up-to-date on testing than other men. The reasons for this are not known but possibilities include that women may be playing a role in encouraging testing for their male partners or that men with unhealthy behaviors are less likely to stay married. Men have previously been found to receive more health benefits from marriage than women, with married men showing lower levels of unhealthy behaviors such as alcohol consumption and cigarette smoking.

This study has several limitations. First, we did not distinguish between the presence of chronic conditions that may increase the chance of cholesterol testing, such as heart disease or diabetes, and those that may distract from testing, such as COPD. Future research that makes this distinction is certainly warranted. However, this limitation does not discount our finding regarding people with no morbidity and no chronic conditions being underscreened. Second, many hospital laboratories in Ontario rely on a global budget and, therefore, do not bill the provincial health insurance plan for their services. Accordingly, we would have missed cholesterol tests that were performed within a hospital setting. The most complex patients may have been most likely to have their cholesterol testing ordered by a hospital-based specialist. Laboratories in rural areas may be more likely to be hospital-based, which may partly explain our result of lower odds of testing for those living in rural versus urban areas; however, results from the regression models were adjusted for rurality. Third, disability and morbidity were measured at the beginning of each person’s 3-year study period and may not have remained consistent over the entire study window. However, our 3-year time frame is still relatively short and it is unlikely that disability and morbidity category designation would change significantly in that time frame for a large number of patients. Fourth, the CCHS is voluntary and self-reported, and respondents may not be representative of the general population. However, it is designed by Statistics Canada to provide reliable population estimates. As well, self-reported morbidity has been found to have predictive validity of disability; measures of self-reported morbidity were strong predictors of disability as determined by medical assessment with tobit regression t-test values ranging from -2.73 to -5.11 in a study by Ferraro and Su. Fifth, it is possible that if we had used a more liberal time frame (ie, 5 years instead of 3 years), then testing gaps would have been less. However, measures of disability and morbidity would have become less accurate and a 3-year period seemed to be the most appropriate based on not knowing participants’ Framingham scores. Sixth, as our outcome is relatively common, the odds ratios described in this study cannot be interpreted as relative risks. Finally, we relied on administrative data that were not expressly collected for...
research purposes and were, therefore, unable to explore such issues as whether patients were offered testing but declined, or whether patients and providers underwent a process of shared decision-making around not testing. It is possible that patients may have had morbidity severe enough to make cholesterol testing inappropriate. However, using these administrative data allowed us to conduct a large, population-based study.

Conclusion

The prevalence of disability, and accordingly the coexistence of disability and multi-morbidity, will rise as our population continues to grow and age. Preventive care may be of particular importance for patients living with disability, but care is not being received equitably. Public health and practice-based interventions focused on reaching those with very low or high health care needs may need to be explored to further address these findings.

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Disclosure

The authors report no conflict of interest in this work.

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