A Likely Magnetic Activity Cycle for the Exoplanet Host M Dwarf GJ 3512

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Abstract

Current radial velocity data from specialized instruments contain a large amount of information that may pass unnoticed if their analysis is not accurate. The joint use of Bayesian inference tools and frequency analysis has been shown as effective in revealing exoplanets but they have been used less frequently to investigate stellar activity. We intend to use radial velocity data of the exoplanet host star GJ 3512 to investigate its magnetic activity. Our study includes the analysis of the photometric data available. The main objectives of our work are to constrain the orbital parameters of the exoplanets in the system, to determine the current level of activity of the star and to derive an activity cycle length for it. An adaptive importance sampling method was used to determine the parameters of the exoplanets orbit. Generalized Lomb–Scargle periodograms were constructed with both radial velocity curve and photometric data. A careful analysis of the harmonic frequencies was conducted in each periodogram. Our fit to multiple Keplerian orbits constrained the orbital parameters of two giant gas planets orbiting the star GJ 3512. The host star showed an increase of its magnetic activity during the last observing campaign. The accurate fit of the radial velocity curve data to the multi-Keplerian orbit permitted to reveal the star rotation in the residuals of the best fit and estimate an activity cycle length of \( \sim 14 \) yr.

Unified Astronomy Thesaurus concepts: Stellar activity (1580); M dwarf stars (982); Exoplanet systems (484); Starspots (1572); Computational methods (1965); Late-type dwarf stars (906); Bayesian statistics (1900); Importance sampling (1892)

1. Introduction

High-precision radial velocity curves of stars are a vast source of information. They may contain signatures of stellar and/or planetary companions, stellar activity, pulsations, and stochastic processes that affect the profile of the spectral lines. Current high-resolution spectrographs attain precisions of \( 1 \) m s\(^{-1}\), well below the detection limit of line profile variations caused by flaring events or by rotating spots (López-Santiago et al. 2003; Berdyugin 2005; Barnes et al. 2011; Flores Soriano & Strassmeier 2017). In addition, high-resolution spectra contain information about the magnetic activity of the star in spectral lines that are formed in the chromosphere. When combined with photometric observations, the study of all these signatures and indicators yields a comprehensive view of the stellar and planetary system.

Different techniques are usually applied to analyze spectroscopic, photometric, and radial velocity data. The most common of those techniques are frequency analysis and model fitting. Frequency analysis is often based on least-squared spectral fitting (Vaníček 1969). Often, frequency analysis comprises the computation of the Lomb–Scargle periodogram (Lomb 1976; Scargle 1982; Zechmeister & Kürster 2009; VanderPlas 2018) and false-alarm probabilities to determine the significance of the peaks revealed (e.g., Balucév 2008). For model fitting, a number of procedures are usually applied, among which the Markov Chain Monte Carlo (MCMC) methods are gaining popularity in all their versions (see Dumasque et al. 2017, for a list of them). MCMC methods are commonly used to approximate the posterior probability density function (pdf) of model parameters, but they demand significant post-processing to derive the so-called Bayesian evidence (a.k.a., model evidence or marginal likelihood). Determining the Bayesian evidence is necessary to discriminate between models (Ford & Gregory 2007; Feroz et al. 2011; Nelson et al. 2020). In particular, for the problem of detecting exoplanets in radial velocity curves, the Bayesian evidence is used to discriminate between models with a different number of planets. Another Bayesian inference method that can be applied to approximate the pdf of the model parameters is importance sampling (IS). The advantage of IS methods is that they yield direct estimates of the Bayesian evidence. IS algorithms were partially applied to exoplanet search problems in the past (Hogg et al. 2010; Loredo et al. 2011; Liu 2014). Commonly, they are combined with MCMC (e.g., Nelson et al. 2016). Regardless of the method used to infer the orbital parameters of exoplanets, only after an accurate model fitting combined with a careful frequency analysis, all the information contained in the radial velocity curve can be extracted.

GJ 3512 is an intermediate-aged mid-M dwarf (Lépine & Gaídos 2011; Terrien et al. 2015) with a moderate magnetic activity. Gizis et al. (2002) detected variable \( \text{H}_\alpha \) emission in four observations of the star \( \langle EW(\text{H}_\alpha) \rangle = -0.548 \) Å; Newton et al. (2017). More recently, Stelzer et al. (2013) determined an upper limit for its X-ray flux \( \log F_X = -13.04 \) erg cm\(^{-2}\) s\(^{-1}\), equivalent to \( L_X \approx 10^{37} \) erg s\(^{-1}\), which is below the mean X-ray luminosity of the Sun during its activity cycle. Finally, Reiners et al. (2018) determined an \( \text{H}_\alpha \) luminosity of \( \log L_{\text{H}_\alpha}/L_{\text{bol}} = -4.4 \) dex, which is similar to that found for other field M dwarfs (López-Santiago et al. 2010; Martínez-Arnaiz et al. 2011). Morales et al. (2019) reported periodic, low-amplitude photometric variability with a period of \( \sim 87 \) days. This value is similar to the rotation period of GJ 3512 determined in other works (e.g., Newton et al. 2016). Morales et al. (2019) also applied a Metropolis-Hastings algorithm to the radial velocity data acquired with the next-generation high-resolution spectrograph CARMENES.
(Quirrenbach et al. 2018). As a result, the authors discovered a giant gas planet with a minimum mass of $\sim 0.46 M_J$ orbiting the star with a period of 203.6 days. The orbit of GJ 3512b is very eccentric ($e > 0.4$), which is difficult to explain without any planet–planet interaction. Therefore, Morales et al. (2019) proposed the existence of a second planet. Their hypothesis was supported also by the sinusoidal shape of the residuals of the radial velocity data after subtraction of the signal of GJ 3512b. A first attempt to derive the orbital parameters of the second planet allowed the authors to determine a lower limit for its period and mass. With a mass of $M > 0.17 M_J$, GJ 3512c appeared to orbit far from the star ($P > 4$ yr).

The aim of the work we present here was to estimate the magnetic activity cycle of the star GJ 3512. For this purpose, a multi-keplerian model was fitted to the radial velocity data to remove the signal of the exoplanets. Then, the residuals were fitted to an amplitude modulated sinusoid as a proxy of the activity cycle of the star. We performed a careful frequency analysis of the spectroscopic and photometric data, as well as the radial velocity curve published in Morales et al. (2019). Different signals in the data were revealed and their nature was investigated. We applied an adaptive IS to infer the parameters of the different models fitted to the data. The Bayesian evidence of each model considered was derived and the most probable configuration for the planetary system was determined. The rotation period of the star and a probable magnetic cycle were revealed only after the subtraction of the signal of the planets. The analysis of contemporary $R$-band photometric observations revealed an increase in the spot activity of the star and suggested the appearance of a flip-flop effect.

2. The Data

The data used for our analysis were published by Morales et al. (2019). The spectroscopic observations were acquired with the high-resolution spectrograph CARMENES (Quirrenbach et al. 2018), mounted on the 3.5 m telescope of the Spanish-German, Calar-Alto Observatory (CAHA) located in Almería (Spain). CARMENES has two channels, referred to as the visual (VIS) and near-infrared (NIR) channels. The former covers the wavelength range 520–960 nm, while the NIR channel covers the range 960–1710 nm. The spectral resolution in the entire spectrum is maintained between 80,000 and 100,000. The instrument configuration permits the detection of chromospheric activity indicators such as the H$_\alpha$ line and the Ca II infrared triplet, if present in the star. Observations were taken from 2016 December to 2019 May. The JD of the first observation is 2457752.4036 days. The details of the observations reduction are given in Morales et al. (2019). The radial velocity measures for GJ 3512 are publicly available at the CDS.\(^3\) We added 10 points at the end of the observations not included by Morales et al. (2019) in their study. These data were provided by the authors in a private communication. Overall, we used 148 measurements from the VIS channel and 149 measures from the NIR channel. The mean error in the VIS channel is 3.94 m s$^{-1}$, while the mean error in the NIR channel is 8.48 m s$^{-1}$.

Several photometric campaigns were carried out by the CARMENES team. For this work, we focused on the Johnson-Cousins $R$-band observations acquired with the robotic 0.8 m Joan Oró telescope at the Montsec Astronomical Observatory (Lleida, Spain). These observations covered a period of approximately 18 months between 2017 December and 2019 May. They were separated into two different seasons, coinciding with the last two seasons of the spectroscopic campaign. The $R$-band observations were acquired with a cadence of a few hours and then combined in a single value for each night, for a total of 161 observations. Although other photometric observations analyzed in Morales et al. (2019) covered the same period, we did not use them to avoid noise caused by the zero-point correction. A wrong zero-point calibration of any instrument could cause uncertainties in the results. Our main goal was to detect photometric variability related to the stellar activity.

3. Frequency Analysis

We started our analysis by investigating the frequency content of the radial velocity curve. A generalized Lomb–Scargle (GLS) periodogram (Zechmeister & Kürster 2009) was computed using the data from both the NIR and the VIS channels. Our aim was to identify significant peaks in the periodogram that could be related to different planets orbiting the star. The GLS periodogram of the data is shown in Figure 1. The $y$-axis represents normalized power. The two dashed lines mark the false-alarm probability of 1% and 0.01%, respectively. These false-alarm probabilities were determined using bootstrap random permutations of the data. The data exhibit three peaks with a high confidence level ($> 99.99\%$). The most intense peak is at a frequency of $4.9 \times 10^{-3}$ Hz ($\sim 202$ days). This is compatible with the results of Morales et al. (2019), who determined a period of $\sim 203.6$ days for GJ 3512b. The second, more intense peak is located at a frequency of $9.6 \times 10^{-3}$ Hz ($\sim 104$ days). This second peak is a harmonic of the feature at $4.9 \times 10^{-3}$ Hz. In fact, several harmonics of this frequency are present in the periodogram. We mark up to six of them in Figure 1. The presence of harmonics in the periodogram of a radial velocity curve indicates that the curve is far from a sinusoid. In other terms, the orbit with the period corresponding to the frequency showing harmonics is highly eccentric. Note that the eccentricity reported by Morales et al. (2019) for the planet discovered by the authors is $e \approx 0.44$.

We were also interested in identifying all the features in the periodogram caused by the signal of GJ 3512b and by the sampling process. Those features may be wrongly interpreted as the signature of other planets. Hence, we constructed a synthetic radial velocity curve with the parameters of GJ 3512b reported by Morales et al. (2019). Then, we determined its periodogram, and compared it with that of the original data. Figure 2 shows that the periodogram of the synthetic curve (red dashed line) accurately fits the periodogram of the data (black continuous line). It includes all the harmonics and the significant peak at approximately $7.5 \times 10^{-3}$ Hz ($\sim 133$ days). The latter is caused by the windowing of the signal. The effect of windowing a time series is well known in signal processing and it is treated in many textbooks (e.g., Prabhu 2018). In practice, the signal is defined in $t \in (−\infty, \infty)$ but it is observed only during a finite time period $t_{\text{obs}} \in [a, b]$. Formally, this amounts to multiplying the observed signal by a square signal. In the frequency domain, this product is translated into the convolution of the Fourier transform of both signals (see Chapter 4 in Oppenheim et al. 1998). The transform of a square signal is a sinc function (e.g., Harris 1978). If the signal of interest is a sinusoid, with a

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\(^3\) http://vizier.u-strasbg.fr/viz-bin/VizieR?source=J/other/Sci/365.1441
Fourier transform consisting in a delta function located at the frequency of the sinusoid, the periodogram of that signal observed during a finite time period is a squared sinc function located at the frequency of the sinusoid (see Figure 5 in VanderPlas 2018). This is the meaning of the peak at $7.5 \times 10^{-3}$ Hz in our periodogram, as it is shown in Figure 2. The lack of uniformity of the sampling mitigates the effect of the remaining peaks of the sinc in the periodogram and they are not clearly observed in Figure 1. Note that, contrarily to the idea of many authors, the less intense peaks created by the windowing are not aliases as they are understood in signal processing. An alias is an artifact caused by the undersampling of the signal.

Figure 2 also shows some disagreement between the two periodograms at very low frequencies (small window in the figure). This difference between the two periodograms suggests the presence of another component in the radial velocity curve. The feature not fitted by the simulated curve is located at $\sim 700$ days. The hypothesis of the second signal is supported by a further analysis of the periodogram. Morales et al. (2019) showed a periodogram of the residuals of the radial velocity curve after subtracting the signal of GJ 3512b. That periodogram is reproduced here in Figure 3. The intense peak is located at a frequency equivalent to a period of approximately 1300 days. The second intense peak in that figure is caused by the data sampling. The authors proposed the existence of a second planet causing that feature. For our work, we did not attempt to construct the periodogram of a simulated curve of this second planet proposed by Morales et al. (2019) because of the lack of constraints in its orbit. The authors could provide
only some lower bounds to the period and the amplitude of the curve for this planet.

According to the above results, we investigated the presence of up to two planets in the radial velocity data. The periods of 202 and 1300 days were used to initialize the proposal pdf in the Bayesian inference algorithm.

4. Bayesian Analysis of the Radial Velocity Curve

4.1. The Method

For the analysis of the radial velocity curve, we used an adaptive importance sampling (AIS) method. In particular, we implemented a version of the algorithm described in Table 2 of Bugallo et al. (2017), with adaptation subject to the quadratic distance between the observations and the model. In particular, the proposal pdf is adapted at some step only if a sample with a lower quadratic distance is found. This distance is used to determine the noise variance that is included in the likelihood function in the next step. In this sense, the scheme acts as a maximum-likelihood algorithm. AIS methods have been used in other research fields but they have been rarely applied by the astrophysical community (Wraith et al. 2009; Lewis & Bridle 2011). Compared to MCMC methods, AIS techniques offer a more natural solution to the problem of model selection because they yield direct estimates of the model evidence without the post-processing needed in MCMC (Stokes et al. 2017; Speagle 2020).

The implementation of the AIS we used for this work is parallelized. This parallelization enables the reduction of the computational run-time and the use of a large number of samples at each step. Note that in most of the MCMC methods commonly applied in the astrophysical literature, one sample is generated at each step (e.g., Gregory 2011; Foreman-Mackey et al. 2013). For our study, we generated $10^8$ samples and performed 1000 iterations. At the end of the process, a total of $10^8$ samples were generated for each configuration, i.e., for the model with a single planet and that of two planets. The Monte Carlo sample that maximizes the likelihood was determined together with the noise variance ($\gamma$ parameter) and the Bayesian evidence. A resampling step was performed at the end of the process to determine confidence levels for each orbital parameter.

4.2. Results of the Inference Method

We searched for the signature of up to two planets in the radial velocity data of GJ 3512. This selection is supported by the frequency analysis of the radial velocity curve (Section 3). Only Keplarian orbits were considered. Commonly, each single orbit is represented by five parameters: period ($P$), amplitude ($K$), eccentricity ($e$), argument of periastron ($\omega$), and last periastron passage ($\tau$). The mean velocity of the star ($v_0$) is a common factor to all the orbits. Therefore, we inverted the following equation:

$$v = v_0 + \sum_{i=1}^{n} K_i \cos(\nu_i + \omega_i + e_i \cos(\omega_i)) + s,$$

where

- $M_i = \frac{2\pi}{P_i}(t - \tau_i),$  \hspace{1cm} (2)
- $M_i = E_i - e_i \sin(E_i),$  \hspace{1cm} (3)
- $\tan \frac{\nu_i}{2} = \frac{1 + e_i}{\sqrt{1 - e_i}} \tan \frac{E_i}{2},$  \hspace{1cm} (4)

and where $n$ is the number of planets we consider and $t$ is the time vector of the observations. $M_i$ is the mean anomaly of the $i$th component, $E_i$ is its eccentric anomaly, and $\nu_i$ is the true anomaly. The vector $s$ is a random variable with zero mean and covariance matrix $\Sigma$ and it represents observational noise. The matrix $\Sigma$ can be expressed as

$$\Sigma = \gamma \Sigma_{\text{obs}},$$  \hspace{1cm} (5)

$$\Sigma_{\text{obs}} = \sigma_{\text{obs}}^2 I \sigma_{\text{obs}}^2,$$  \hspace{1cm} (6)

Figure 3. GLS periodogram of the radial velocity curve residuals after subtracting the curve of GJ 3512b. The intense peak at $\sim1300$ days is marked. The second intense peak is caused by the sampling scheme.
Here, $I$ is the identity matrix, $\sigma_{\text{obs}}$ is the standard deviation of the radial velocity determined at $t$, $\sigma_{\text{obs}}$ is the transpose of the standard deviation, and $\gamma$ is the variance of the noise to be inferred by optimization (see Section 4.1). The scalar $\gamma^{1/2}$ is commonly interpreted as a jitter induced by the stellar magnetic activity (e.g., Oshagh et al. 2017). However, it contains also information of the model truncation.

Prior and proposal pdfs were defined for each parameter. Uniform density functions were selected for the prior pdfs. For the proposal pdfs, we considered uniform density functions only for $e$ and $\omega$, while we used normal pdfs for the remaining parameters. The choice of normal proposal pdfs permits easy adaptation in the algorithm. The mean and variance of those normal pdfs were chosen according to the data. For instance, the semi-amplitude of the radial velocity curve is $70 \pm 20$ m s$^{-1}$ and, thus, the mean of the normal pdf for $K$ was chosen inside that range. The standard deviation for the same parameter was chosen large enough to permit the algorithm to jump to other regions of the parameter space. Similar arguments are applied to the remaining proposal pdfs. The algorithm is adaptive and will explore the parameter space conveniently. Note that we did not fix the periods selected from the frequency analysis (Section 3). Windowing also causes leakage (Harris 1978; Phillips et al. 2008), which shifts the peaks in the periodogram if the time window does not cover exactly a multiple of the period of the signal. Therefore, the peaks in a periodogram are not always located at the actual frequency of the signal. AIS algorithms adapt the proposal pdf at each step. In this work, the proposal pdfs were adapted according to the Monte Carlo sample that maximizes the likelihood. Table 1 lists the prior and initial proposal pdfs for the two components considered in this work. Note that we adapted only the Gaussian proposals. The uniform proposals are preserved at each step.

The Bayesian (or model) evidence $Z$ was determined for each model. We calculated $Z$ ratios to settle the most probable number of planets. The results are shown in Table 2. The model with one planet (hereafter M1) is clearly less probable than the model with two planets (M2). Note that the ratios in Table 2 are in logarithmic scale. The presence of the less massive planet is also supported by the value of $\gamma_{\text{M2}}$, which is lower than $\gamma_{\text{M1}}$. Another indicator for model selection is the Bayesian information criterion (BIC) of Schwarz (1978). The BIC is a logarithmic magnitude that accounts for the increase in the number of parameters. It is defined as

$$\text{BIC} = k \ln n - 2 \ln \hat{L},$$

where $k$ is the number of dimensions (the number of parameters of the model to be estimated), $n$ is the number of observations (data points), and $\hat{L}$ is the maximum value of the likelihood function for the candidate model. If the model errors are assumed independent and normally distributed, the BIC can be determined as

$$\text{BIC} = n \ln \hat{\sigma}^2 + k \ln n,$$

which indicates that M2 is much more probable than M1 (the model with the lowest BIC value is preferred).

The orbital parameters we inferred for both models are listed in Table 3, together with the masses of the planets (we assumed $M_\odot = 0.123 M_\odot$; Morales et al. 2019). The best fit (M2) is plotted in Figure 4. For any of the models (M1 and M2), the orbit and the mass of GJ 3512b determined in this work are essentially the same as those obtained by Morales et al. (2019).

In model M2, the orbital parameters inferred for GJ 3512c are in agreement with the ranges proposed by Morales et al. (2019) for their suggested second planet. We inferred a period of $P \approx 1600$ days, a semi-amplitude of $K = 14.1$ m s$^{-1}$, and an eccentricity of $e \approx 0$ for this planet. The minimum mass derived for GJ 3512c, $m \sin i = 0.20 M_\odot$, is in the range of the super-Neptune planets. This configuration perfectly reproduces the observed periodogram. Figure 5 shows the periodogram of the data (continuous, black line) and that for the simulated curve of the two planets. Contrarily to the simulated curve for a single planet, this configuration is able to reproduce the periodogram at low frequencies (small window in the two figures). As a result, the presence of other planets in the system cannot be inferred from the frequency analysis of the current data. We notice that the small
uncertainty in the period of GJ 3512c is a product of the joint posterior distribution of these data, which is very narrow. The AIS algorithm draws many samples and it assigns weights to all them. Then, it adapts the mean values and the variances of the normal proposal pdfs according to the weighted means and variances of the drawn samples and repeat the process until the requested number of iterations is completed. At the end of the process, hundreds of millions of samples along with their corresponding weights are generated. A resampling process is performed with those weights to determine the posterior distribution. For this data set, the marginal posterior distribution of the period of GJ 3512c is very narrow. Nevertheless, other observations could give slightly different results. Note that this procedure is different from the MCMC methods, where all the accepted samples are used to construct the marginal posterior distributions without weighting. This result is intuitive in some sense and also extremely important, in our opinion. We believe that this point deserves further studies. Indeed, without considering this result, one may believe that the chain produced by an MCMC algorithm or the weighted samples generated by an IS scheme fully represent the posterior pdf but, actually, they just represent the tails of the posterior distribution (i.e., meaningless regions). Namely, it is extremely important to discover the narrow region (a peak) of very high values of the posterior.

The process of Bayesian inference was repeated with the measures from the VIS and NIR channels separately, for the model M2. The value determined for each orbital parameter and each planet is shown in Table 4. The results obtained previously for the entire VIS+NIR data set are included in this table for comparison. Table 4 shows that the inference method obtains very similar results for the orbital parameters of GJ 3512b with any of the three combinations: VIS, NIR, and VIS+NIR. Likewise, the orbital parameters of GJ 3512c inferred with the VIS channel data are very similar to those determined for the entire data set. Contrarily, the value of the period of GJ 3512c inferred with the NIR data is very different. Nevertheless, this value is very uncertain. This result shows that the marginal posterior pdf has a strong dependence on the data, especially if the period of the planet is not fully covered.

5. Magnetic Activity of GJ 3512

5.1. Rotation, Spots, and the Flip-flop Effect

GJ 3512 is a moderately active M5.5 dwarf. Morales et al. (2019) determined a rotation period of 87 ± 5 days. The authors found peaks in the periodograms of photometric data of

![Figure 4: Radial velocity curve of GJ 3512. The black continuous line is the best fit to the two-planet model (M2) with AIS. The red dashed line is the signal of the first planet. The blue dotted line is the second planet. The gray curves are states that maximize the likelihood at each step. The gold, filled circles are observations from the VIS channel. The green, filled squares are observations from the NIR channel. In the bottom panel, the residuals of the fits are plotted.](image-url)
the star at different bands consists with previous determinations by Newton et al. (2016). However, signatures of this rotation period are not clearly detected in the magnetic activity indices. Figure 6 shows the GLS periodograms for those indices (Hα line width, differential line width, and chromatic index) compared to the periodogram of the R-band photometric data. The dashed line delimits the false-alarm probability of 0.01%. Only the photometric data show clearly the signature caused by the rotation period of the star. In the R band, the peak in the periodogram is observed at ~86 days. In the periodogram of the Hα line and differential line profile measures, a feature can be observed around that frequency but with very low power. This result is not surprising because the Hα line width of M dwarfs is a tracer of the stochastic activity of the star (e.g., Crespo-Chacón et al. 2006).

To produce the periodogram of the R-band photometric observations in Figure 6, we used all the available data. An offset was applied to the values of the first data set like in Morales et al. (2019). This offset compensates zero-point differences between the two TJO observation campaigns. The value of the offset was determined by comparing the means of the observations in the first observation period to those of the second observation period. The corrected light curve is shown in Figure 7. The first set of observations shows a decrease in the magnitude of the star with a duration of approximately 47 days. After that, it remains nearly flat. Contrarily, the R-band photometric data acquired later shows more variability. The period of photometric observations coincides with the last two sets of radial velocity observations (see Figure 4). Therefore, we constructed a new GLS periodogram with only R-band photometric data from the last data set, which expands for approximately 300 days. The result is shown in Figure 8. In this periodogram, two intense signals are detected. The signal at a lower frequency corresponds to a period of 86 days. The second, and more intense, peak is the first harmonic of the previous signal and it is located at a period of ~42 days. These two signals were noticed by Morales et al. (2019, supplementary material) but the authors obtained lower significance possibly due to the mixture of data from different instruments. The frequency ratio of both peaks is 2.02. The presence of harmonics in the periodogram of photometric data from stars was investigated by Reinhold & Arlt (2015). The authors simulated the light curves of stars with a different number of spots at distinct latitudes and obtained some relations between the heights of the peaks of the harmonics and the latitude of the spots. In particular, they concluded that low-latitude spots cause intense peaks of the harmonics in the periodogram and less sine-shaped light curves. In Figure 8, the intensities of the fundamental period and the first harmonic are similar and the light curve is far from a sine shape. If the argument of Reinhold

Table 4

| Parameter | VIS Channel | NIR Channel | VIS + NIR Channels |
|-----------|-------------|-------------|--------------------|
| v₀ (m s⁻¹) |            |             |                    |
| P (day)   | GJ 3512b: 203.90±0.02 | 1594.3±3.1 | 1395.5±10.5 |
|           | GJ 3512c: 0.23         | 15.6±0.4   | 14.5±1.1  |
| K (m s⁻¹) | GJ 3512b: 72.73±0.18 | 73.69±1.1 | 71.8±0.2  |
|           | GJ 3512c: 15.9±0.4    | 14.5±1.1  | 15.1±0.2  |
| e         | GJ 3512b: 0.43±0.02  | 0.45±0.01  | 0.44±0.01  |
|           | GJ 3512c: 0.0043±0.0009 | 0.0078±0.0004 | 0.0183±0.0001 |
| ω (rad)   | GJ 3512b: 2.182±0.025 | 2.183±0.027 | 2.195±0.006 |
|           | GJ 3512c: 0.0007±0.0002 | 0.0008±0.0013 | 0.0012±0.0001 |
| τ¹ (day)  | GJ 3512b: 196.3±0.1 | 197.1±1.4   | 196.5±0.5  |
|           | GJ 3512c: 1016.5±0.7 | 942.6±191.8 | 1119.2±3.4 |

Note. Last periastron passage prior to the first observation.
& Arlt (2015) is valid, GJ 3512 would contain a large spot at a very low latitude.

Although the hypothesis of the low-latitude single spot is attractive, a very intense first harmonic peak as that detected by us was not derived from the simulations of Reinhold & Arlt (2015). Neither did Santos et al. (2017) obtain very intense first harmonics in their simulations of single spots at distinct latitudes. High first harmonic intensities like that observed in GJ 3512 were detected by Balona (2013) in several A-type stars. The author attributed the intensity of the harmonics to the
Another plausible explanation for the intense first harmonic in the periodogram of GJ 3512 is the presence of two spots at longitudes separated by $\sim 180^\circ$. A similar configuration was observed in the M4 dwarf GJ 1243 (Davenport et al. 2015). In that star, the primary spot was stable over years, while the secondary spot decayed in 100–500 days. These timescales are compatible with the observations of GJ 3512, where the fundamental period of 84–90 days seems stable but the first harmonic is detected only in the last data set. The photometric curve of GJ 3512 is also compatible with the observation of a flip-flop effect (Berdyugina 2005; Hackman et al. 2013). This effect consists in the inversion in longitude of the location of the spots on the star. During the inversion process, spots of similar size coexist at longitudes separated by approximately 180° (e.g., Berdyugina et al. 1999). This pattern may produce the first harmonic in the periodogram like in GJ 3512 (Figure 8).

5.2. Activity Cycle and Rotation Period in the Velocity Curve

GJ 3512 shows an emission H$_\alpha$ line (Gizis et al. 2002), with $\log L_{H_\alpha}/L_{bol} = -4.4$ dex (Reiners et al. 2018). The latter is common to many old M dwarfs (López-Santiago et al. 2010). No magnetic activity saturation is expected for this star (see Figure 7 in Martínez-Arnáiz et al. 2011). Therefore, it is likely that GJ 3512 has a long activity cycle similar to that of the Sun. Magnetic activity cycles have been observed in several low-mass stars. Lorente & Montesinos (2005) studied the relation between the cycle length and the rotation period of a number of active stars. According to their results, GJ 3512 should have a magnetic activity cycle of $P_{cycle} > 20$ yr. However, Böhm-Vitense (2007) showed that inactive stars have shorter activity cycles than those observed in active stars with the same rotation rate. From that study, a cycle length of approximately 15 yr can be inferred for GJ 3512. This value is similar to the activity period found for GJ 551, another M5.5 inactive dwarf (Suárez Mascareño et al. 2016).

In view of the previous discussion, we investigated the presence of the magnetic activity cycle in the radial velocity data. The signature of the 87 day period is not observed in the periodogram of the radial velocity curve (Figure 1), but it is common that a signal passes unnoticed in a periodogram if it is not persistent in time. Note that the power of the signal generated by GJ 3512b is very high and it can mask the remaining signals (Section 3). Therefore, an amplitude modulated sinusoid was fitted to the residuals using AIS. Our aim was to find the magnetic activity cycle of GJ 3512. The
A model used was

\[ v = a \sin \left( \frac{2\pi}{P_{\text{cycle}}} t + \phi_1 \right) \sin \left( \frac{2\pi}{P_{\text{rot}}} t + \phi_2 \right) + b, \]

(10)

where \( a \) is the amplitude of the sinusoid, \( P_{\text{cycle}} \) is the period of the activity cycle, \( P_{\text{rot}} \) is the rotation period, \( \phi_1 \) and \( \phi_2 \) are phases, and \( b \) is an offset (for a similar analysis using solar data, see Edmonds 2016). The results are shown in Table 3 and the best fit is plotted in Figure 9. The residuals of the fit have a standard deviation of 5.66 m s\(^{-1}\). This value is close to the mean uncertainty of the VIS channel and it is below the mean uncertainty of the NIR channel. It is also lower than the standard deviation of the residuals from the two-planet scheme that enables the determination of the actual intensity of the observational noise. As a result, the residuals of the best fit contain only information on the stellar activity and observational noise.

6. Summary and Conclusions

We have presented a study of the exoplanetary system GJ 3512 that comprises stellar activity and planetary orbits analysis. We found that the radial velocity curve contains information about the magnetic activity of the star that was revealed only after an accurate fit of the curve to a multi-Keplerian orbit. For the fitting, we applied a Bayesian inference method named AIS. Our algorithm contains an optimization scheme that enables the determination of the actual intensity of the observational noise. As a result, the residuals of the best fit contain only information on the stellar activity and observational noise.

The two planets orbiting GJ 3512 are giant gas planets with minimum masses of 0.5 and 0.2 \( M_J \) respectively. Both planets are located far from the habitability zone of the star. The orbit of GJ 3512b is highly eccentric. This fact was already noticed by Morales et al. (2016). Contrarily, the orbit of GJ 3512c is nearly circular. The high eccentricity of the inner planet is not unusual, but the presence of the external giant gas planet could have stabilized its orbit (see Morales et al. 2019). It seems unlikely that a terrestrial planet exists in the habitable zone due to the current configuration of the planetary system. Inner planets have been found in systems with an external high-eccentric perturber (e.g., Otor et al. 2016), but we did not detect any signal of such internal planet in the periodogram of the residuals.

The accurate fit of the exoplanet’s orbit to the data of GJ 3512 revealed the signature of the star rotation period in the residuals of the best fit. The sinusoidal shape of the curve is more evident during the third observational period. This result agrees with the aspect of the photometric curve. The frequency analysis of the \( R \)-band observations during the same time interval showed two intense features. The first signature corresponds to the star rotation period. The second and more intense feature is the first harmonic of this rotation period. The first harmonic of the star rotation period is observed also in the periodogram of the radial velocity curve residuals. We suggest that two active regions showed dark spots at opposite longitudes during that period.

We further fitted the residuals of the two-planet model fit to the radial velocity curve using an amplitude modulated (AM) sinusoid. The residuals of this fit to an AM sinusoid show low standard deviations, close to the mean uncertainty of the VIS channel measures. The carrier signal has a period of 88 days, in agreement with the rotation period of GJ 3512. The amplitude of the carrier signal is likely modulated with a period of \(~14\) yr. We interpreted this modulation as the activity cycle length of the star. This value agrees with the activity cycle of slow rotators found in the literature. We notice that the value of the cycle length is not well constrained. Further observations will be needed to determine a more accurate cycle length for GJ 3512.
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References
Balona, L. A. 2013, 
Balona, L. A., Catanzaro, G., Abedigamba, O. P., Ripepi, V., & Smalley, B. 2015, 
Balouc, R. V. 2008, 
Barnes, J. R., Jeffers, S. V., & Jones, H. R. A. 2011, 
Berdyugina, S. V. 2005, 
Berdyugina, S. V., Berdyugin, A. V., Ilyin, I., & Tuominen, I. 1999, 
Böhm-Vitense, E. 2007, 
Bugallo, M. F., Elvira, V., Martino, L., et al. 2017, 
Crespo-Chacón, I., Montes, D., García-Alvarez, D., et al. 2006, 
Davenport, J. R. A., Hebb, L., & Hawley, S. L. 2015, 
Dumusque, X., Borsa, F., Damasso, M., et al. 2017, 
Edmonds, I. 2016, arXiv:1605.05028
Feroz, F., Balan, S. T., & Hobson, M. P. 2011, 
Flores Soriano, M., & Strassmeier, K. G. 2017, 
Ford, E. B., & Gregory, P. C. 2007, in ASP Conf. Ser., 371, Statistical Challenges in Modern Astronomy IV, ed. G. J. Babu & E. D. Feigelson (San Francisco, CA: ASP), 189.
Foreman-Mackey, D., Hogg, D. W., Lang, D., & Goodman, J. 2013, 
Gizis, J. E., Reid, I. N., & Hawley, S. L. 2002, 
Gregory, P. C. 2011, 
Hackman, T., Pelt, J., Mantere, M. J., et al. 2013, 
Harris, F. J. 1978, Proc. IEEE, 66, 51
Hogg, D. W., Myers, A. D., & Bovy, J. 2010, 
Lépine, S., & Gaidos, E. 2011, 
Lewis, A., & Bridle, S. 2011, 
Liu, B. 2014, 
Lomb, N. R. 1976, 
López-Santiago, J., Montes, D., Fernández-Figueroa, M. J., & Ramsey, L. W. 2003, 
López-Santiago, J., Montes, D., Gálvez-Ortiz, M. C., et al. 2010, 
Loredo, T. J., Berger, J. O., Chernoff, D. F., Clyde, M. A., & Liu, B. 2011, 
Lorente, R., & Montesinos, B. 2005, 
Martínez-Arnáiz, R., López-Santiago, J., Crespo-Chacón, I., & Montes, D. 2011, 
Morales, J. C., Mustill, A. J., Ribas, I., et al. 2019, 
Nelson, B. E., Ford, E. B., Buchner, J., et al. 2020, 
Nelson, B. E., Robertson, P. M., Payne, M. J., et al. 2016, 
Newton, E. R., Irwin, J., Charbonneau, D., et al. 2016, 
Newton, E. R., Irwin, J., Charbonneau, D., et al. 2017, 
Oppenheim, A., Willsky, A., Nawab, S., & Hernández, G. 1998, 
Oshagh, M., Santos, N. C., Figueira, P., et al. 2017, 
Otor, O. J., Montet, B. T., Johnson, J. A., et al. 2016, 
Phillips, C., Parr, J., & Riskin, E. 2008, 
Prabhu, K. 2018, 
Quirrenbach, A., Amado, P. J., Ribas, I., et al. 2018, 
Reiners, A., Zechmeister, M., Caballero, J. A., et al. 2018, 
Reinhold, T., & Arlt, R. 2015, 
Santos, A. R. G., Cunha, M. S., Avelino, P. P., García, R. A., & Mathur, S. 2017, 
Santos, A. R. G., Cunha, M. S., Avelino, P. P., García, R. A., & Mathur, S. 2017, 
Scargle, J. D. 1982, 
Schwarz, G. 1978, 
Speagle, J. S. 2020, 
Stelzer, B., Marino, A., Micela, G., López-Santiago, J., & Liebke, C. 2013, 
Stokes, B., Tylu, F., & Hudson, I. 2017, in AIP Conf. Ser., 1853, American Institute of Physics Conference Series (Melville, NY: AIP), 11003
Suárez Mascareño, A., Rebolo, R., & González Hernández, J. I. 2016, 
Terrien, R. C., Mahadevan, S., Bender, C. F., Deshpande, R., & Robertson, P. 2015, 
Vaníček, P. 1969, 
Wraith, D., Kilbinger, M., Benabed, K., et al. 2009, 
Zechmeister, M., & Kürster, M. 2009, 

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