Short Communication

Are diet–prostate cancer associations mediated by the IGF axis? A cross-sectional analysis of diet, IGF-I and IGFBP-3 in healthy middle-aged men

D Gunnell, SE Oliver, TJ Peters, JL Donovan, R Persad, M Maynard, D Gillatt, A Pearce, FC Hamdy, DE Neal and JMP Holly

High circulating levels of insulin-like growth factor-I (IGF-I) are associated with an increased risk of developing prostate cancer (Chan et al., 1998; Harman et al., 2000; Stattin et al., 2000; Chokkalingam et al., 2001) and other cancers (Holly et al., 1999). In serum and body fluids, IGF-I’s activity is regulated by a complex system of six binding proteins and an acid-labile subunit. Most (90%) circulating IGF-I is bound to IGF binding protein-3 (IGFBP-3) and associations of IGFs with prostate cancer are generally strongest with the molar ratio (IGF-I/IGFBP-3) or in statistical models controlling for IGFBP-3 (Chan et al., 1998; Harman et al., 2000; Stattin et al., 2000; Chokkalingam et al., 2001). Raised levels of bioavailable IGF-I may, therefore, increase cancer risk, and raised IGFBP-3, by reducing IGF-I’s bioavailability, may reduce risk.

IGF-I plays a role in energy and protein metabolism as well as modulating cell turnover and apoptosis (Thissen et al., 1994; Holly et al., 1999). Energy restriction leads to reduced production of IGF-I (Thissen et al., 1994), and animal experiments suggest that this pathway mediates the cancer-protective role of diet restriction (Dunn et al., 1997). Other dietary influences on IGFs may underlie some of the diet–prostate cancer associations observed. While no specific food or nutrient is an established risk factor for prostate cancer, dietary aspects most consistently related to its risk are red meat, animal fat, calcium and dairy product consumption and lower consumption of vegetables (Kolonel, 1996; World Cancer Research Fund, 1997; Department of Health, 1998). Diets rich in tomatoes, a major source of the carotenoid lycopene, are associated with reduced risk (Giovannucci, 1999).

Several, generally small, cross-sectional studies have examined the association of diet with the IGF axis (Darling-Raedke et al., 1998; Kaklamani et al., 1999; Allen et al., 2000; Signorello et al., 2000; Mucci et al., 2001; Holmes et al., 2002; Giovannucci et al., 2003). The largest investigation (n = 1037) (Holmes et al., 2002) reported that higher levels of energy, protein and milk intake were associated with raised IGF-I and high fat intake with low IGFBP-3, broadly consistent with previous, smaller, studies (Kaklamani et al., 1999; Ma et al., 2001). Other studies have reported reduced levels of IGF-I with tomato consumption (Mucci et al., 2001) and vegan diets (Allen et al., 2000). There has been only one (Giovannucci et al., 2003) large-scale investigation of the association of diet with IGF in community-based men.

MATERIALS AND METHODS

Within a case–control study nested in a population-based investigation into the early detection and management of prostate cancer (ProtecT; Donovan et al., 2002), stored blood samples from 368 disease-free men (controls) were assayed for IGF-I and IGFBP-3. Controls were matched to cases on age, general practice and date of recruitment. Included in this analysis are the 344 (95%) of these disease-free men who completed a 114-item validated food-frequency questionnaire (FFQ) (Bingham et al., 1997). Over two-thirds (n = 242) of the men also provided information on occupation, smoking and physical activity and had height and weight measured. Ethical approval was obtained from the relevant multicentre and local research ethics committees.
Based on FFQ responses, and using standard food tables (The Royal Society of Chemistry and MAFF, 1991) and portion size data for men of this age (Ministry of Agriculture and Food, 1993), we estimated weekly consumption of: energy, carbohydrate, protein, total fat, saturated and polyunsaturated fat, calcium, red meat, dairy products, vegetables, milk, tomatoes and foods containing tomatoes (baked beans, tomato ketchup and tomato juice). These were selected on the basis of research findings and reviews examining associations of diet with prostate cancer (World Cancer Research Fund, 1997; Giovannucci, 1999; Chan and Giovannucci, 2001) and the IGF-axis (Ma et al, 2001; Holmes et al, 2002). Given difficulties in measuring lycopene from FFQs (Kristal and Cohen, 2000), we used the frequency of reported consumption of tomatoes and products with high tomato content.

**Laboratory methods**

Non-fasted blood specimens, taken using standard techniques, were spun and frozen to −80°C within 18 h. For the IGF-I assays, an ELISA kit was used (Diagnostic Systems Laboratories, TX, USA). Assays for serum IGFBP-3 used a previously validated ‘in-house’ double antibody radioimmunoassay (Cheetham et al, 1998). The average coefficients of variation for intra-assay variability for IGF-I and IGFBP-3 were 3 and 3.6%, and for inter-assay variation were 15 and 14%. To measure (crudely) bioavailable IGF-I, we multiplied the molar ratio of IGF-I/IGFBP-3 by 5.33 (molecular weights 40,000 and 7500 Da, respectively).

**Statistical analysis**

Using Stata (Stata Corporation, 2001) we calculated age-, centre- and energy-adjusted levels of IGF-I, IGFBP-3 and the molar ratio in quartiles of the distribution of each dietary factor. Intakes of individual food groups were considered in three *a priori* categories. Adjustment for overall energy intake means that dietary measures relate to dietary composition rather than absolute intake (Willett, 1998).

Least-squares linear regression models investigated change in growth factor levels for a one standard deviation increase in each dietary factor. Log or square-root transformations were used for the latter due to positive skewness, and sampling weights adjusted for the dependence on the age distribution of cases. Tests for trend were based on the continuous variable (for nutrients and food groups) or three-level category (for tomato products and milk).

We assessed possible confounding by exercise, smoking, body mass index (BMI) and socioeconomic position in the 242 men with complete data.

**Table 1**  Age-, centre- and energy-adjusted levels of IGF-I, IGFBP-3 and IGF-I/IGFBP-3 molar ratio in relation to quartiles of increasing intake of dietary variables (*n* = 344)*a*  

|          | Quartile 1 (low intake) | 2 | 3 | 4 (high intake) | Change in growth factor (95% CI) per s.d. increase in each dietary variable | *P*-value (linear trend)*b* |
|----------|-------------------------|---|---|-----------------|--------------------------------------------------------------------------------|-----------------------------|
| IGF-I (ng ml⁻¹) |                        |   |   |                 |                                                                                |                             |
| Energy intake (MJ) | 119.9                  | 133.8 | 129.3 | 125.8 | 1.2 (−3.6 to 5.9) | 0.63 |
| Carbohydrates (g) | 109.2                  | 128.6 | 129.6 | 141.5 | 10.0 (−12.2 to 21.2) | 0.08 |
| Protein (g)      | 115.7                  | 137.4 | 131.5 | 124.9 | 2.6 (−6.5 to 11.8) | 0.57 |
| Fat (g)          | 121.1                  | 125.4 | 125.3 | 137.2 | 5.0 (−4.6 to 14.6) | 0.30 |
| Polyunsaturated fat (g) | 114.4             | 122.6 | 136.2 | 136.1 | 9.1 (1.6 to 16.5) | 0.017 |
| Saturated fat (g) | 125.3                  | 121.4 | 128.5 | 133.9 | 1.8 (−6.5 to 10.2) | 0.67 |
| Red meat (g)     | 125.2                  | 134.2 | 135.3 | 117.8 | −2.7 (−8.3 to 2.9) | 0.35 |
| Dairy products (g) | 117.8                 | 130.7 | 125.6 | 134.1 | 4.4 (−0.8 to 9.7) | 0.09 |
| Calcium (mg)     | 120.5                  | 119.1 | 137.2 | 131.8 | 6.5 (0.5 to 12.5) | 0.035 |
| Vegetables (g)   | 134.1                  | 125.9 | 122.8 | 126.5 | −2.1 (−8.3 to 4.1) | 0.50 |
| IGFBP-3 (ng ml⁻¹) |                        |   |   |                 |                                                                                |                             |
| Energy intake (MJ) | 3281.6                 | 3479.5 | 3494.2 | 3328.1 | 15.4 (−10.1 to 32.4) | 0.80 |
| Carbohydrates (g) | 3017.2                 | 3423.7 | 3541.0 | 3598.2 | 184.5 (−131.6 to 500.7) | 0.25 |
| Protein (g)      | 3120.7                 | 3488.3 | 3651.0 | 3395.1 | 150.0 (−107.2 to 407.2) | 0.25 |
| Fat (g)          | 3564.5                 | 3387.7 | 3299.6 | 3359.9 | −85.8 (−329.8 to 158.3) | 0.49 |
| Polyunsaturated fat (g) | 3272.0             | 3288.7 | 3357.0 | 3667.2 | 228.1 (−1.2 to 457.4) | 0.05 |
| Saturated fat (g) | 3465.4                 | 3514.0 | 3444.0 | 3162.4 | −170.4 (−347.7 to 339) | 0.10 |
| Red meat (g)     | 3469.4                 | 3389.2 | 3521.4 | 3196.7 | −78.0 (−196.5 to 40.4) | 0.20 |
| Dairy products (g) | 3348.8                 | 3450.2 | 3385.1 | 3396.2 | 3.9 (−139.4 to 147.3) | 0.96 |
| Calcium (mg)     | 3345.5                 | 3326.7 | 3540.2 | 3368.1 | 81.2 (−256.1 to 218.5) | 0.37 |
| Vegetables (g)   | 3164.8                 | 3637.4 | 3401.5 | 3374.8 | 93.8 (−357.7 to 243.4) | 0.22 |
| IGF-I:IGFBP-3 molar ratio |                |   |   |                 |                                                                                |                             |
| Energy intake (MJ) | 0.206                  | 0.221 | 0.210 | 0.213 | −0.001 (−0.012 to 0.010) | 0.85 |
| Carbohydrates (g) | 0.202                  | 0.214 | 0.209 | 0.226 | −0.009 (−0.034 to 0.016) | 0.48 |
| Protein (g)      | 0.204                  | 0.231 | 0.204 | 0.213 | −0.002 (−0.021 to 0.018) | 0.88 |
| Fat (g)          | 0.193                  | 0.208 | 0.214 | 0.235 | 0.011 (0.007 to 0.029) | 0.23 |
| Polyunsaturated fat (g) | 0.200               | 0.212 | 0.231 | 0.208 | −0.002 (−0.017 to 0.013) | 0.84 |
| Saturated fat (g) | 0.208                  | 0.196 | 0.210 | 0.238 | 0.012 (−0.006 to 0.030) | 0.19 |
| Red meat (g)     | 0.204                  | 0.224 | 0.221 | 0.209 | 0.001 (−0.008 to 0.011) | 0.82 |
| Dairy products (g) | 0.205                  | 0.209 | 0.212 | 0.225 | 0.007 (−0.003 to 0.018) | 0.18 |
| Calcium (mg)     | 0.207                  | 0.204 | 0.218 | 0.222 | 0.006 (−0.008 to 0.019) | 0.39 |
| Vegetables (g)   | 0.236                  | 0.206 | 0.201 | 0.209 | −0.011 (−0.022 to −0.000) | 0.045 |

*aAll values are controlled for age, study centre and energy intake, except that for energy intake, which is controlled for age and study centre only. All models are weighted by the inverse of the sampling probability in relation to age. *bTests for trend based on continuous measure of diet.
RESULTS

Mean age was 62.2 years (range 50–70) and most men (90%) were nonsmokers and came from nonmanual social classes (64%). Mean (s.d.) blood levels of IGF-I, IGFBP-3 and the molar ratio (IGF-I/IGFBP-3) were 126.6 ng ml$^{-1}$ (36.9), 3393.6 ng ml$^{-1}$ (1049.8) and 0.21 (0.08), respectively. Median daily intakes were as follows: energy: 10.3 MJ; carbohydrate: 314.8 g; protein: 89.5 g; fat: 77.6 g; red meat: 46.9 g; dairy products 344.9 g; calcium 1126.3 g; vegetables 271.0 g.

Raised IGF-I levels were seen in men consuming higher levels of polyunsaturated fat ($P_{\text{trend}} = 0.017$) and calcium ($P_{\text{trend}} = 0.035$) (Table 1). There were weaker positive associations with carbohydrate and dairy products. IGFBP-3 levels were weakly positively associated with polyunsaturated fats ($P_{\text{trend}} = 0.05$) and inversely associated with saturated fats ($P_{\text{trend}} = 0.10$). The molar ratio was inversely related to vegetable intake ($P_{\text{trend}} = 0.045$).

Controlling for BMI, social class, smoking and exercise attenuated the associations of IGF-I and IGFBP-3 with carbohydrates, polyunsaturated fats and, to a lesser extent, vegetables (not shown). Associations of dairy products and calcium with IGF-I and IGFBP-3 were not confounded.

IGF-I tended to be lower and IGFBP-3 higher in those who ate tomatoes or tomato-containing products more frequently, although evidence for a trend was only clear for IGF-I/IGFBP-3 molar ratio (Table 2). Men consuming higher levels of milk had raised levels of IGF-I ($P_{\text{trend}} = 0.004$). There was no association between milk intake and IGFBP-3 and the molar ratio was highest in those men drinking at least one pint of milk per day. These associations were little changed after adjustment for BMI, social class, smoking and exercise (not shown). Associations of IGF-I with milk were not confounded by calcium intake, whereas associations with calcium intake were attenuated in models controlling for milk intake.

Excluding men ($n = 95$) reporting low levels of energy intake in relation to their estimated basal metabolic rate (ratio of energy intake/basal metabolic rate <1.2 (Joint FAO/WHO/UN Expert Consultation, 1985)) did not change the associations with milk intake, tomato-rich products or vegetable intake.

DISCUSSION

In a group of healthy, community-sampled, middle-aged men, we found associations of the IGF-axis with several aspects of diet linked previously to prostate cancer. Positive relations were seen with dairy products, milk and calcium intake, all of which were associated with raised IGF-I levels. High intakes of vegetables and

Table 2 Age-, centre- and energy-adjusted mean levels of IGF-I, IGFBP-3 and IGF-I/IGFBP-3 molar ratio in relation to increasing levels of intake of tomatoes, tomato-rich products and milk*
tomatoes or tomato-containing products were associated with lower levels of IGF-I or its molar ratio. In contrast to some other studies (Kaklamani et al, 1999; Holmes et al, 2002), we found only weak associations with saturated fat and no evidence of an association with red meat.

Associations were not confounded by socioeconomic position or lifestyle. While we have examined associations with a range of dietary variables and three different measures of the IGF-axis, thereby increasing the possibility of chance results, our findings are consistent with previous studies. In a cross-sectional study, it is not possible to determine whether dietary associations arise as the result of long-term intake of particular foods/nutrients or reflect patterns of intake around the time of blood sampling. Association of elevated IGF-I levels with dietary products, milk and calcium are consistent with some (Heaney et al, 1999; Ma et al, 2001; Holmes et al, 2002) but not all (Mucci et al, 2001) previous analyses. The strongest evidence of a causal association between higher levels of milk consumption and IGF-I comes from a randomised trial of dietary milk supplementation, reporting a rise in IGF-I in those supplemented but not the controls (Heaney et al, 1999). While some research suggests that neonates absorb IGF-I from breast milk (Diaz-Gomez et al, 1997), there is no strong evidence that bovine IGF-I in cows milk could be similarly absorbed from the gut (Holmes et al, 2002). Dietary intake of animal protein (essential amino acids) is known to stimulate IGF-I production (Thissen et al, 1994), but we found no evidence of associations with animal protein intake, nor that controlling for animal protein intake attenuated associations with milk (not shown). This contrasts with the findings of Giovannucci et al (2003) and Holmes et al (2002). In Giovannucci et al’s analysis, associations with vegetable protein were, however, of similar magnitude to those for animal protein.

Associations of calcium, milk and dairy products with IGF-I suggest a possible pathway linking dietary intake of these factors with prostate cancer (Chan and Giovannucci, 2001). The relation of these dietary aspects with prostate cancer risk are, however, in the opposite direction to their association with colorectal cancer (Ma et al, 2001; Wu et al, 2002), another neoplasm associated with raised IGF-I levels (Ma et al, 1999). The IGF–cancer associations seen for a range of different cancer sites (Holly et al, 1999; Yu and Rohan, 2000) may not therefore be explained in terms of common dietary influences on the growth factor axis. Nevertheless, our finding that vegetable intake was weakly related to lower molar ratios is consistent with the observation that vegetable-rich diets appear to protect against colorectal, breast and prostate cancer (World Cancer Research Fund, 1997), although associations of vegetable intake with IGF-I or IGFBP-3 have not been found in other studies (Kaklamani et al, 1999; Holmes et al, 2002).

The weak associations of IGFs with tomatoes and tomato-containing products support those reported for 112 Greek men (Mucci et al, 2001), where a strong inverse association was found between cooked tomato consumption and IGF-I. Likewise, in the Nurses Study intake of lycopene was positively associated with circulating levels of IGFBP-3 (but not IGF-I). These findings hint at the possible importance of the IGF axis in mediating the protective effect of higher levels of tomato or lycopene intake on prostate cancer reported in several investigations (Giovannucci, 1999). A possible biological mechanism lies in the reported inhibitory effects of lycopene on IGF-I receptor signalling and cell cycle progression (Karas et al, 2000), but a small trial of lycopene supplementation found no difference in IGF-I levels in supplemented vs control subjects (Kucuk et al, 2001).

International comparisons of cancer incidence and changes in incidence in migrants moving between different continents, indicates large dietary influences on epithelial cancer incidence (World Cancer Research Fund, 1997). Our study adds to evidence that aspects of diet previously linked to prostate cancer may influence cancer risk through the IGF-axis. Trials of dietary interventions aimed at reducing bioavailable IGF-I are now required. Identification of relevant aspects of diet could then lead to trials of dietary interventions against cancer incorporating measurements of IGF-I.

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