1. Introduction

Detached eclipsing binary stars with well-determined parameters are one of the best methods for testing stellar evolutionary models (Lastennet & Valls-Gabaud 2002; Torres et al. 2010). These models play an important role in areas, such as exoplanet research, where they are used to determine the mass and age of planet-host stars, and in galaxy evolution modelling, where they are used to predict the properties of the stars that are observed by surveys of the Milky Way and other galaxies (Schaerer 2013).

Stellar evolutionary models can use different prescriptions to describe complex phenomena such as convection and diffusion. The free parameters used in these prescriptions are, in general, poorly constrained by the underlying physics, so they are either calibrated from observations, or by detailed simulations. Creating models for stars that are different from the stars used in the calibrations, whilst using the same calibrated parameters values, can introduce systematic errors into the models. As such, these free parameters are still large sources of uncertainty in using stellar models to infer the ages and other properties of stars (Lebreton et al. 2014). Recently, research has been pushing to understand these parameters in more detail, with Valle et al. (2017) looking at how uncertainties on observational data affect the determination of these parameters, in particular the overshooting parameter. They note, following on from an earlier paper (Valle et al. 2016), that for binaries where both stars are on the main-sequence, the calibration of the overshooting parameter is hampered by uncertainties of the order 0.5% in the radii and 1% in the masses. This indicates that in order to calibrate the overshooting, binary systems with stars at two different evolutionary stages must be used and their parameters must be known to very high precision, (at least 1% in the masses, Valle et al. 2017). The number of well-studied binary systems with both the required precision, and the required evolutionary state are very few. Excluding Al Phe and TZ For (see below) there is only one other system in the list by Torres et al. (2010) that meets these requirements, and can be classed as a solar-type star, that is V432 Aur. This system also has a relatively short 3-day orbital period, meaning there is a strong possibility that tidal interactions between the stars would invalidate the assumption that the two stars have evolved independently. Some examples which meet the requirements include LL Aqr (Graczyk et al. 2016), TZ For (Valle et al. 2017), and two binaries in NGC 6791 (Brogard et al. 2012). Each have provided a constraint on the helium abundance of the particular binary system. Kirkby-Kent et al. (2016) looked at the binary system Al Phe, and after

WASP 0639-32: a new F-type subgiant/K-type main-sequence detached eclipsing binary from the WASP project

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ABSTRACT

Aims. Our aim is to measure the masses and radii of the stars in a newly-discovered detached eclipsing binary system to a high precision (&1%), enabling the system to be used for the calibration of free parameters in stellar evolutionary models.

Methods. Photometry from the Wide Angle Search for Planets (WASP) project was used to identify ISWASP J063930.33-322404.8 (TYC 7091-888-1, WASP0369-32 hereafter) as a detached eclipsing binary system with total eclipses and an orbital period of $P = 11.66$ days. Lightcurve parameters for WASP0639-32 are obtained using the EBOP lightcurve model, with standard errors evaluated using a prayer-bead algorithm. Radial velocities were measured from 11 high-resolution spectra using a broadening function approach, and an orbit was fitted using SBOP. Observed spectra were disentangled and an equivalent width fitting method was used to obtain effective temperatures and metallicities for both stars. A Bayesian framework is used to explore a grid of stellar evolution models, where both helium abundance and mixing length are free to vary, and use observed parameters (mass, density, temperature, and metallicity) for each star to obtain the age and constrain the helium abundance of the system.

Results. The masses and radii are found to be $M_1 = 1.1544 \pm 0.0043 M_{\odot}$, $R_1 = 1.833 \pm 0.023 R_{\odot}$, and $M_2 = 0.7833 \pm 0.0028 M_{\odot}$, $R_2 = 0.7286 \pm 0.0081 R_{\odot}$ for the primary and secondary, respectively. The effective temperatures were found to be $T_1 = 5400 \pm 7833$ K and $T_2 = 5800 \pm 80$ K for the primary and secondary star, respectively. The system has an age of $4.2^{+1.0}_{-1.3}$ Gyr, and a helium abundance in the range $0.251$–$0.271$.

Conclusions. WASP0639-32 is a rare example of a well-characterised detached eclipsing binary system containing a star near the main-sequence turn-off point. This makes it possible to measure a precise age for the stars in this binary system and to estimate their helium abundance. Further work is needed to explore how this helium abundance estimate depends on other free parameters in the stellar models.

Key words. binaries: eclipsing – stars: solar-type – stars: fundamental parameters – stars: evolution
improving the mass and radius estimates of both stars, looked at how independently varying the mixing length and helium abundance affected the age of the system. Again the precise observational constraints allowed the helium abundance $Y_1$ to be constrained to the range 0.25–0.28.

If meaningful statistics are to be done on the helium abundance in low-mass stars, then more systems need to be studied to the same level of precision. This paper details the photometric and spectroscopic analysis of a newly-discovered detached eclipsing binary system containing an evolved F-type primary and a main-sequence K-type star, which was first identified using data from the Wide Angle Search for Planets (WASP; Pollacco et al. 2006). Section 2 describes the data that used in the photometric analysis (Sect. 3) and the spectra used in the orbital analysis (Sect. 4) and the spectroscopic analysis (Sect. 5). Section 6 combines the lightcurve and spectroscopic orbit parameters to give masses and radii of the system and in Sect. 7 we investigate how the helium abundance can be constrained by the precise binary parameters system. Finally, discussion and conclusion sections follow in Sects. 8 and 9.

2. Observations

2.1. WASP photometry

The WASP project (Pollacco et al. 2006) uses two instruments, one located at the Observatorio del Roque de los Muchachos, La Palma, and the other at Sutherland Observatory, South Africa. Each instrument consists of eight wide-field cameras each with a 2048 × 2048 pixel CCD. In total, 39,621 photometric measurements were obtained for WASP 0639-32 (1SWASP J063930.33-322404.8, TYC 7091-888-1) between May 2006 and March 2012 by the WASP-South instrument using 200 mm, f/1.8 lenses and broadband 400–700 nm filters (Pollacco et al. 2006). A specialised pipeline (Pollacco et al. 2006) is used to reduce the images and the data are subject to a detrending algorithm described by Collier Cameron et al. (2006). The WASP data can suffer from significant amounts of scatter caused by clouds, scattered moonlight, etc. To remove the affected observations, the photometric data is process by an algorithm, that was first described in Kirkby-Kent et al. (2016). We briefly discuss the process here.

Firstly, each measurement that has the weighting factor, $\sigma_{\text{err}}$, set to zero, is removed. This factor is defined as part of the WASP pipeline ($\sigma_{\text{err}}$) in Collier Cameron et al. (2006) and characterises the amount of scatter present from external sources. When set to zero, it indicates the pipeline as has marked the value as missing or bad. Data where the uncertainty in a flux measurement was more then five times the median value were also removed. One WASP field contained only ten observations, which were removed as they caused problems when the analysis compared to the expected value within in a phase bin. If observations differed by more than ten times the uncertainties for each phase bin, it was removed, and if more than 80% of data from a block was offset, the entire block was removed.

The final step looks for data that is significantly offset from the other data. This technique uses a model generated from 1000 bins across the phase-folded data, with the median magnitude in each bin used as a reference. The magnitude from each observation within a block of data (the observation from one camera on one night) was compared to the expected value within in a phase bin. If observations differed by more than ten times the uncertainties for each phase bin, it was removed, and if more than 80% of data from a block was offset, the entire block was removed.

Overall, from the initial 39,621 observations stored in the WASP archive, 28,566 remained for use during the analysis. Table 1 provides a summary of the number of points removed during this initial processing stage.

| Reason          | Points removed |
|-----------------|----------------|
| $\sigma_{\text{err}} = 0$ | 4870           |
| $f_{\text{err}} > 5\times\text{median}$ | 2408           |
| Small field     | 10             |
| Offset          | 3766           |
| Remaining       | 28,566         |

Notes. Last row states the number of observations remaining.

The eleven spectra used in this work, were obtained using the Ultraviolet and Visual Échelle Spectrograph (UVES; Dekker et al. 2000) at the Very Large Telescope (VLT), over a period between 9th October 2014 and 20th January 2015. A wavelength range of approximately 500–700 nm, with a small gap at 600 nm, and a signal-to-noise ratio of over 100 was obtained for each spectrum. The slit was set to a width of 0.7 arcseconds, resulting in spectra with a resolving power $R = 56990$. Each spectrum was reduced using the standard UVES reduction pipeline (Ballester et al. 2000).

3. Photometric Analysis

3.1. Ephemeris

To find the period, $P$, and time of primary minimum, the WASP photometry, with very minimal cleaning applied (i.e. only the points with $\sigma_{\text{err}} = 0$ where removed), was fitted using JKTEBOP (Southworth et al. 2004). The surface brightness ratio, sum of radii, ratio of radii, orbital inclination, and $e \cos \omega$ and $e \sin \omega$ parameters were included in the model fitting. The results for the period and time of primary minimum are shown in Eq. (2).

$$HJD \text{ Pri. Min.} = 2455 \pm 0.33931(26) + 11.658317(5) \ E.$$ \ (2)

The period and time of primary minimum have been kept fixed in all subsequent analysis. Figure 1 shows the photometric measurements phase-folded using the ephemeris in Eq. (2), after the processing described in Sect. 2.1.

3.2. Lightcurve Modelling

To model the lightcurve we used the same method as described in Kirkby-Kent et al. (2016). It uses the subroutine $\text{light}$ from the EBOP lightcurve analysis code (Nelson & Davis 1972; Popper & Etzel 1981), with modifications allowing it to be
called directly from the Python programming language. The parameters used in the EBOP model are optimised using the least-squares Levenberg–Marquardt Python module, MPFIT (Markwardt 2009).

The model has seven parameters that were fitted: surface brightness ratio at the centre of the stellar discs, $J$; sum of the radii, $r_{\text{sum}} = r_1 + r_2$; ratio of the radii, $k = r_2/r_1$; inclination, $i$; $e \cos \omega$, $e \sin \omega$ and third-light, $l_3$. In addition to these fitted parameters, the fractional radii, $r_1$ and $r_2$, are automatically calculated from $r_{\text{sum}}$ and $k$, while the eccentricity, $e$, and longitude of periapsis, $\omega$, are calculated from $e \cos \omega$ and $e \sin \omega$.

The mass ratio, $q = M_2/M_1$, was fixed at 0.67. This value was determined by initially setting $q = 0.5$, fitting a model and then using the estimates of $r_1$, $r_2$, and the spectroscopic orbit parameters (see Sect. 4) as inputs into JKTBDSDIM\(^1\). For WASP 0639-32, the mass ratio contributes very little to the overall shape of the lightcurves meaning variations of 0.01 will not alter the final best-fit parameters. The gravity darkening exponents were fixed at 0.26 and 0.46 for the primary and secondary component, respectively (Claret & Bloemen 2011). These parameters also have very little effect on the lightcurve as the stars are nearly spherical in shape.

Attempts to include the limb darkening coefficients as free parameters in the fitting proved unsuccessful as these coefficients are not well-constrained by the data. Instead, the coefficients were fixed at values taken from interpolating between values in the tables by Claret & Bloemen (2011). The Kepler passband was used to approximate the response of the WASP broadband filter. The adopted primary limb-darkening coefficient was $u_p = 0.50 \pm 0.05$ and $u_r = 0.63 \pm 0.03$ was adopted for the secondary. To account for the uncertainties in the limb-darkening coefficients, models were calculated with the fitted limb darkening parameters varied by their uncertainties. The average scatter in the parameters from these models has been added in quadrature to the uncertainties from the best-fit model. Table 2 details the uncertainty in each parameter due to the uncertainty in the limb-darkening coefficients.

As in Kirkby-Kent et al. (2016), the parameter-space is explored using Markov chain Monte Carlo (MCMC), in the form of the Python module, emcee (Foreman-Mackey et al. 2013). It was used to ensure the lightcurve solution was not a local minimum in the parameter space. emcee uses an affine-invariant ensemble sampling (stretch-move) algorithm of Goodman & Weare (2010). By using affine-invariant transformations, the algorithm can work with skewed distributions. The algorithm explores the parameter space using a group of walkers, which can be split allowing the process to run in parallel. Walker

\[ \ln L(y; \Theta) = -\frac{1}{2} \sum_{n=1}^{N} \left[ \left( \frac{m_n - y_n(\Theta)}{m_{\text{err},n}} \right)^2 - \ln \left( \frac{2\pi}{m_{\text{err},n}} \right) \right], \]

where $y$ is a vector of length $N$ containing the magnitudes generated for a model, $\Theta$ is a vector containing the varying parameters ($J$, $r_{\text{sum}}$, $k$, $i$, $e \cos \omega$, $e \sin \omega$, and $l_3$), $m$ is the observed magnitude and $m_{\text{err}}$ is the standard error on the magnitude. Priors were applied, but these are only used to prevent the parameter exploring areas that are unphysical, e.g., $r_{\text{sum}}$ or $J$ being less than zero. The MCMC process ran using 150 walkers for 2500 steps, of which the first 200 were discarded to allow for an adequate burn-in stage. For each of the walkers, a starting point for each parameter was chosen by choosing a number at random from a normal distribution (with a mean of zero and variance of 0.01) and adding it to the best-fit parameter. This method creates a “ball” of walkers close to the best-fit solution; however, during the burn-in stage, the walkers spread out from this ball. The chains for each parameter were checked to ensure the burn-in stage was completed within these first 200 steps and adequate mixing had occurred. We also check for bias in the starting positions by carrying out a run where the walkers’ starting positions can be distributed by up to three times the uncertainty on the parameters, as determined by the covariance matrix obtained from the best fit. The burn-in stage for this test was longer, but still produced the same distribution. Various checks were performed to ensure suitable mixing and convergence. The walkers’ positions were plotted against step number to visually check the mixing and to determine the burn-in length. Convergence is checked through a running-mean test. We also check the acceptance fraction (mean = 0.43) and auto-correlation length (max = 92) for each run. The auto-correlation length is a measure of the number of evaluations needed to have independent samples (Foreman-Mackey et al. 2013). Parameter uncertainties were calculated using a prayer-bead method, again as described in Kirkby-Kent et al. (2016). The prayer-bead algorithm used 500 shifts, spread evenly across the data, with starting positions selected randomly from the MCMC analysis. The prayer-bead algorithm allows robust estimates of the uncertainties associated with lightcurve parameters (as the standard deviation of the best-fit values from the 500 shifts) and

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**Table 2. Uncertainty contribution to each parameter from uncertainty in the limb darkening coefficients.**

| Parameter | Original | Detrend |
|-----------|----------|---------|
| $J$       | 0.0042   | 0.0042  |
| $r_{\text{sum}}$ | 0.00006 | 0.00003 |
| $k$       | 0.003    | 0.003   |
| $i$ ($'$) | 0.005    | 0.005   |
| $e \cos \omega$ | 0.000003 | 0.000005 |
| $e \sin \omega$ | 0.002   | 0.002   |
| $l_3$     | 0.005    | 0.005   |
| $r_1$     | 0.00013  | 0.00013 |
| $r_2$     | 0.0013   | 0.0013  |
| $\omega$ ($'$) | 0.0020 | 0.0019  |

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\(^1\) [http://www.astro.keele.ac.uk/jkt/codes/jktabsdim.html](http://www.astro.keele.ac.uk/jkt/codes/jktabsdim.html)
accounts for the correlated noise that can be found within the WASP data, and varying the starting positions overcomes any bias that may be present by the choice of initial parameters. Figure 2\(^\text{5}\) shows the density distribution for \(r_1\) and \(r_2\), which has been calculated from the \(k\) and \(r_{\text{sum}}\) positions explored during the MCMC. The cut-off in the distribution occurs where the third-light becomes zero, as there are strong correlations between \(k\) and \(l_j\). The parameters \(k\) and \(r_{\text{sum}}\) are directly related to observable features in the lightcurve and so there values are well determined.

### 3.3. Detrending Investigations

To help ensure the most accurate parameters possible were obtained from the lightcurve, we have investigated whether the WASP detrending algorithm alters the best-fit parameters for WASP0639-32. The algorithm locates and corrects for four trends of systematic errors found across all stars in one field. It can be described using

\[
\tilde{m}_{i,j} = m_{i,j} - \sum_{k=1}^{M} c_j^{(k)} a_i, \quad (4)
\]

where \(m_{i,j}\) and \(\tilde{m}_{i,j}\) are the observed and corrected magnitude, respectively, for star \(j\) at time \(i\). \(M\) is the total number of trends, \(a_i\) are basis functions detailing the patterns of systematic errors and \(c_j\) describes to what extent each basis function affects a particular star.

Using the method described in Kirkby-Kent et al. (2016), effective detrending coefficients, \(c'\), were calculated using singular value decomposition (SVD). The effective detrending coefficients include a fixed binary lightcurve model, \(L\), when they are calculated to allow the variability of an eclipsing to be taken into consideration. It can be described using

\[
\tilde{m}_i = m_i + L_i - \sum_{k=1}^{M} c_j^{(k)} a_i, \quad (5)
\]

where \(m_i\) and \(\tilde{m}_i\) are the observed and corrected magnitude (respectively), and \((k)a_i\) are the same detrending basis functions as in Eq. (4). Once calculated, the trends are removed from the observed data and a new model fitted. Initially the best-fit model from Sect. 3.2 was used for \(L\), but the coefficients were calculated in an iterative process with slightly different models until each parameter change by less than 0.005\% from the previous model. As before, the same MCMC and prayer-bead analysis was used as in Sect. 3.2.

Table 3 contains best-fit parameters for both the original model fit and those from the detrending testing. The difference between the two is also shown for ease of comparison. In all cases the difference is less than the uncertainties associated with that parameters, so the WASP detrending algorithm does not affect the lightcurve parameters for WASP0639-32. This is the same conclusion as was found the analysis of AI Phe (Kirkby-Kent et al. 2016). The detrended parameters have been used in the remaining analysis. Figure 3 shows the

![Fig. 2. Density distribution of the fractional radii, \(r_1\) and \(r_2\), for WASP0639-32. Grey cross marks the values calculated from the best-fit values from the fit to the “original” data.](image)

![Fig. 3. Upper panels: detrended best-fit model for WASP0639-32 (grey line) plotted over the WASP-South photometry for the primary (left) and secondary (right) eclipses. Lower panels: residuals, with the grey line marking zero.](image)

| Parameter | Original | Detrended | Difference |
|-----------|----------|-----------|------------|
| \(J\)     | 0.4527(72) | 0.4513(71) | 0.0014     |
| \(r_{\text{sum}}\) | 0.0946(42) | 0.0949(42) | -0.00034   |
| \(k\)     | 0.3949(53) | 0.3975(53) | -0.0026    |
| \(i\)     | 89.9925(68) | 89.9943(67) | -0.0018    |
| \(e\cos\omega\) | -0.00027(12) | -0.00026(12) | -0.00001   |
| \(e\sin\omega\) | -0.0077(44) | -0.0070(44) | -0.0007    |
| \(l_3\)   | 0.0077(20) | 0.012(20) | -0.0045    |
| \(r_1\)   | 0.06785(86) | 0.06797(86) | -0.00012   |
| \(r_2\)   | 0.02679(29) | 0.02720(30) | -0.00023   |
| \(e\)     | 0.0077(54) | 0.0070(53) | 0.0007     |
| \(\omega\) | 268.0(4.8) | 267.9(4.9) | 0.1        |

Notes. Standard errors on the final two digits of each parameter value are given in the parentheses and include the contribution from the uncertainties in the limb darkening co-efficients.
the system is very close circular. In circular systems degeneracy from Eq. (2) because the fitted lightcurve parameters showed that periastron were fixed at the period and time of primary minimum. The orbital parameters. The following parameters were included (Foreman-Mackey et al. 2013) to ensure robust uncertainties on Kₜ and Bₒ, and systematic uncertainties between the two stars. This method was adapted to include the semi-amplitude of the primary and secondary, K₁ and K₂, respectively. The orbital period and time of periastron were fixed at the period and time of primary minimum from Eq. (2) because the fitted lightcurve parameters showed that the system is very close circular. In circular systems degeneracy means it is not possible to fit the eccentricity, angle of periastron and time of periastron simultaneously. Initial fits showed an average offset of 0.56 km s⁻¹ between the observed and computed (O–C) measurements for the secondary star. As such, a constant offset parameter, Bₒ, was fitted alongside the orbital parameters. It is likely that this offset is due the large difference in the spectral types of the two stars, having focused the template on the hotter, primary star. The offset will also, in part, be due the effects of differing gravitational redshifts and convective blueshifts between the two stars (Lindegren & Dravins 2003).

For one component, j, the log-likelihood can be written as

\[
\ln L_j(y_{rv}; \Theta_{sys}) = -\frac{1}{2} \sum_{n=1}^{N} \left( \frac{y_{rv,n} - y_{model,n}(\Theta_{sys})}{s_{j,n}} \right)^2 - \ln \left( \frac{2\pi}{s_{j,n}} \right),
\]

where \(y_{rv,j}\) is a vector of length N containing the modelled radial velocities of star j, \(\Theta_{sys}\) is a vector containing the varying parameters (e, \(\omega\), \(\gamma_{sys}\), K₁, K₂, Bₒ, \(\sigma_{sys,1}\), \(\sigma_{sys,2}\)) and \(r_j\) are measured radial velocities for component j. \(s_{j,n}\) are weights for each measurement that combine the measured internal uncertainties, \(\sigma_{j,n}^2\) and a systematic uncertainty \(\sigma_{sys,j}^2\), such that

\[
s_{j,n}^{-2} = s_{j,n}^{-2} + \sigma_{sys,j}^{-2},
\]

\(s_{j,n}^2\) represents a combination of instrumental uncertainty and stellar jitter for component j. Stellar jitter is more significant for more evolved stars (Wright 2005), therefore \(\sigma_{sys,j}^2\) was included separately for each component. The log-likelihood from both components are summed together as \(\ln \mathcal{L} = \ln \mathcal{L}_1 + \ln \mathcal{L}_2\) where the labels \(j = 1, 2\) denote the primary and secondary component, respectively. The chosen priors place loose constraints on the chain parameters to prevent the walkers exploring regions of the parameter-space that correspond to non-physical values. Prior on the semi-amplitude velocities are set to allow exploration between −500 and 500 km s⁻¹, and the systemic velocity is restricted to between −200 and 500 km s⁻¹. The MCMC used 500 walkers and ran for 1000 steps, with the first 400 steps removed as burn-in steps. The chains from each parameter were checked via a running-mean to ensure the chains had converged, and the walkers’ positions were plotted against step number to ensure there is suitable mixing. The acceptance fraction (mean=0.40) and parameter auto-correlation lengths (max=87) are also checked.

The resulting orbit parameters are shown in Table 5 and the resulting model orbits are plotted in Fig. 4. The parameters and their associated uncertainties are calculated from the 15.9, 50, and 84.1 percentiles. The eccentricity is consistent with a circular orbit, but it has been allowed to vary in order for the uncertainty to be accounted for in K₁ and K₂. K₁ and K₂ are the key parameters used to determine the masses of the stars. Both are determined directly from the range of the radial velocity curve, so their values are well determined.

### 4. Spectroscopic Orbit

Radial velocity measurements for both stars were made using the broadening function method (Rucinski 1992, 2002; Rucinski & Lu 1999) in RaveSpan (Pilecki et al. 2017). The broadening function smoothing was set to a value of 3 and a template spectrum with parameters of \(T_{eff} = 6200K\), \(\log g = 4.0\), [Fe/H]=0.0, and \(v\sin(i) = 0.0\), km s⁻¹ was used. The measured radial velocities are shown in Table 4. The spectra were obtained on separate nights to avoid correlations between measured radial velocities.

Orbital parameters were fitted using SBOP, the Spectroscopic Binary Orbit Program (written by P. B. Etzel), with an MCMC wrapper in the form of the Python module emcee (Foreman-Mackey et al. 2013) to ensure robust uncertainties on the orbital parameters. For a detached eclipsing binary system, the log-likelihood can be written as

\[
\ln L_j(y_{rv}; \Theta_{sys}) = -\frac{1}{2} \sum_{n=1}^{N} \left( \frac{y_{rv,n} - y_{model,n}(\Theta_{sys})}{s_{j,n}} \right)^2 - \ln \left( \frac{2\pi}{s_{j,n}} \right),
\]

where \(y_{rv,j}\) is a vector of length N containing the modelled radial velocities of star j, \(\Theta_{sys}\) is a vector containing the varying parameters (e, \(\omega\), \(\gamma_{sys}\), K₁, K₂, Bₒ, \(\sigma_{sys,1}\), \(\sigma_{sys,2}\)) and \(r_j\) are measured radial velocities for component j. \(s_{j,n}\) are weights for each measurement that combine the measured internal uncertainties, \(\sigma_{j,n}^2\) and a systematic uncertainty \(\sigma_{sys,j}^2\), such that

\[
s_{j,n}^{-2} = s_{j,n}^{-2} + \sigma_{sys,j}^{-2},
\]

\(s_{j,n}^2\) represents a combination of instrumental uncertainty and stellar jitter for component j. Stellar jitter is more significant for more evolved stars (Wright 2005), therefore \(\sigma_{sys,j}^2\) was included separately for each component. The log-likelihood from both components are summed together as \(\ln \mathcal{L} = \ln \mathcal{L}_1 + \ln \mathcal{L}_2\) where the labels \(j = 1, 2\) denote the primary and secondary component, respectively. The chosen priors place loose constraints on the chain parameters to prevent the walkers exploring regions of the parameter-space that correspond to non-physical values. Prior on the semi-amplitude velocities are set to allow exploration between −500 and 500 km s⁻¹, and the systemic velocity is restricted to between −200 and 500 km s⁻¹. The MCMC used 300 walkers and ran for 1000 steps, with the first 400 steps removed as burn-in steps. The chains from each parameter were checked via a running-mean to ensure the chains had converged, and the walkers’ positions were plotted against step number to ensure there is suitable mixing. The acceptance fraction (mean=0.40) and parameter auto-correlation lengths (max=87) are also checked.

The resulting orbit parameters are shown in Table 5 and the resulting model orbits are plotted in Fig. 4. The parameters and their associated uncertainties are calculated from the 15.9, 50, and 84.1 percentiles. The eccentricity is consistent with a circular orbit, but it has been allowed to vary in order for the uncertainty to be accounted for in K₁ and K₂. K₁ and K₂ are the key parameters used to determine the masses of the stars. Both are determined directly from the range of the radial velocity curve, so their values are well determined.

### 5. Spectroscopic Analysis

#### 5.1. Disentangling

We have used our own implementation of the matrix disentangling algorithm of Simon & Sturm (1994) to obtain individual spectra of the two stars. This method was adapted to include the apparent flux ratio between the stars for each observed spectrum as additional parameters in the disentangling. This allows the spectrum taken during the secondary eclipse to be included with a flux ratio of zero in the disentangling analysis. It is assumed that all spectra obtained out of eclipse have the same flux ratio. A simple grid search algorithm was used to optimise this parameter by finding the value of this flux ratio that minimises the

| Parameter | Best-fit value | Parameter | Best-fit value |
|-----------|----------------|-----------|----------------|
| K₁ (km s⁻¹) | 47.32(8) | \(\omega\) (°) | 269.96(12) |
| K₂ (km s⁻¹) | 69.74(11) | Bₒ (km s⁻¹) | 0.85(13) |
| \(\gamma_{sys}\) (km s⁻¹) | 57.46(8) | \(\sigma_{sys,1}\) (km s⁻¹) | 0.26(7) |
| \(e\) | 0.0009(1.2) | \(\sigma_{sys,2}\) (km s⁻¹) | 0.09(10) |

Table 4. Radial velocities for WASP 0639-32 measured using broadening function method within RaveSpan.

| HJD−2 450 000 | Primary (km s⁻¹) | Secondary (km s⁻¹) |
|---------------|------------------|--------------------|
| 6939.78595    | 100.90           | −5.61              |
| 6939.85977    | 101.44           | −6.86              |
| 6940.81320    | 104.62           | −10.61             |
| 6956.80762    | 20.29            | 113.83             |
| 6958.81865    | 14.74            | 122.02             |
| 6985.78700    | 92.35            | 7.10               |
| 6999.65212    | 100.42           | −4.35              |
| 6999.77011    | 98.29            | −3.15              |
| 7015.77003    | 12.02            | 124.57             |
| 7040.65621    | 16.67            | 117.65             |
| 7042.59910    | 58.74            | −                  |

Table 5. Fitted spectroscopic orbit parameters for WASP 0639-32.
root mean square (RMS) residuals between the observed spectra and the spectra reconstructed from the disentangled spectra. This search algorithm forms part of a Python wrapper that converts the UVES spectra into the appropriate format for disentangling algorithm. First, the wrapper re-interpolates the spectra onto a uniform logarithmic wavelength scale in the barycentric reference frame. A median filter is used as an initial normalisation, points affected by cosmic rays are replaced with a median value, and the spectrum is binned to provide the a sampling similar to the original UVES spectra. The disentangling is focused on small sections of spectrum, 30 Å in length, around specific Fe I and Fe II lines. By choosing a small region around each line, uncertainties from the continuum placement are reduced, compared with a fitting a continuum to the entire spectrum. The processing is done for each observed spectrum, and is done independently for each small wavelength section. Sections of spectrum were processed around each iron line present in the line-list of Doyle et al. (2013). Residuals between the observed spectra and reconstructed spectra were used to search for and remove remaining cosmic rays using three times the standard deviation of the residuals as a rejection criterion. The affected regions are replaced with the fitted values. The spectra were normalised and the disentangling algorithm is run one final time, using the optimal flux ratio. Figure 5 shows an example of the resulting disentangled spectra around the Fe I at 543.452 nm, where \( L_2/L_1 = 0.0513 \).

Figure 6 compares the observed and recombined disentangled spectra for four of the eleven observed spectra, around the same Fe I line that is shown in Fig. 5. The residuals, which have been scaled up by a factor of ten so they can be seen, are shown above the spectra. There are no strong features in the residuals. Large spikes near the absorption lines would indicate the radial velocities are incorrect, or large slopes at the end of a segment in the continuum of Fig. 5 would suggest there was an issue with the normalisation. The four spectra shown in Fig. 6 are a representative sample of the set.

5.2. Equivalent width fitting

An automated method is used to obtain a table of equivalent widths (EWs) for selected Fe line of each star, independently. Functions from the 2014 version of iSpec (Blanco-Cuaresma et al. 2014a) are used to first normalise and fit a continuum to each segment of disentangled spectrum. A small region around the expected location of the iron line is searched to identify the line, and cross-match the atomic data with the information in the Doyle et al. (2013) line-list. In some cases multiple matches were found, in which case the one closest the expected wavelength was chosen. Once a line has been identified, a Gaussian profile is fitted to the line and is used to determine the EW for a particular line. Each fit has been visually checked to ensure a sensible fit was made. Any lines that were blended or could be significantly offset from the continuum were removed from the final selection EWs. In total, there are 38 EWs for the primary (28 Fe I lines and 10 Fe II lines) and 19 EWs for the secondary star (16 Fe I lines and 3 Fe II lines).

An EW fitting procedure, wrapped in the MCMC Python module, emcee (Foreman-Mackey et al. 2013), was used to determine the best effective temperature, surface gravity, metallicity and microturbulence, \( (T_{\text{eff}}, \log g, [\text{Fe/H}], \text{v}_{\text{mac}}, \text{respectively}) \) for each star separately. Initial walker positions are generated by perturbing initial parameter estimates by multiplying numbers drawn randomly from a normal distribution by step parameters \((100, 0.05, 0.1, \text{and } 0.05 \text{ for } T_{\text{eff}}, \log g, [\text{Fe/H}], \text{and } \text{v}_{\text{mac}}, \text{respectively})\). The step parameters are needed because of the different scales associated with the different parameters. For a given set of parameters, \( \Omega_{\text{fit}} = [T_{\text{eff}}, \log g, [\text{Fe/H}], \text{v}_{\text{mac}}, \text{respectively}] \), small sections of synthetic spectrum are generated around the lines used for the EW measurements using solar abundances from Asplund et al. (2009), MARCS.GES (Gustafsson et al. 2008) model atmospheres and the line-list and atomic data of Doyle et al. (2013). The projected rotational velocity, \( v \sin i \), and macroturbulence, \( v_{\text{mac}} \), are not fitted, and the iSpec’s default parameters are used \((v \sin i = 2 \text{ km s}^{-1} \text{ and } v_{\text{mac}} = 3 \text{ km s}^{-1})\) when generating the sections of synthetic spectrum. These parameters are not consider in the MCMC fitting because they do not alter the measured EW (Niemczura et al. 2014). Once measured, the EWs from the synthetic spectrum sections \( W_{s} \) are compared to the observed
EWs, $W_o$, to judge the choice of $\Theta_{\text{sp}}$. The overall log-likelihood function can be written as

$$
\ln L(W; \Theta_{\text{sp}}) = -\frac{1}{2} \sum_{i=1}^{N} \left( \frac{W_{o,i} - W_{s,i}(\Theta_{\text{sp}})}{\sigma_{W,i}} \right)^2 - \ln \left( \frac{2\pi}{\sigma_{W,i}} \right), \quad (8)
$$

where $W$ is a vector of length $N$ containing the fitted EWs, and $\sigma_{W,i}$ are the uncertainties associated with the measured EWs. Priorors were used to prevent the walkers exploring outside of the limits set by the model atmospheres, i.e., $4500 < T_{\text{eff}} < 7500$, $3.5 < \log g < 5.0$, $-1.0 < [\text{Fe/H}] < 1.0$, and $0 < v_{\text{mic}} < 100$. For both the primary and secondary star, 100 walkers and 500 steps were used. For both stars, the first 100 steps were removed to allow for adequate burn-in. Auto-correlation lengths are between 50 and 70 for the different runs, and acceptance fractions are 0.34–0.55.

Table 6 shows the resulting spectroscopic parameters for both stars in WASP 0639-32, for two different cases. Parameter uncertainties are calculated from the 15.9, 50, and 84.1 percentiles. Where a single standard error is quoted, we have used the larger of the similar values for the positive and negative error bars. The first case, labelled “free”, has the surface gravity included as a free parameter in the EW-fitting. In the second case, labelled “fixed”, the surface gravity was fixed at the values shown in Table 6. For both the primary and secondary star, 100 walkers and 500 steps were used. For both stars, the first 100 steps were removed to allow for adequate burn-in. Auto-correlation lengths are between 50 and 70 for the different runs, and acceptance fractions are 0.34–0.55.

![Fig. 6](image-url) Observed (black) and recomputed disentangled spectra (grey dashed) around the Fe I line at 543.452 nm for four spectra from WASP 0639-32. The residuals, scaled by a factor of $10$, are shown 0.25 above each spectrum segment. Plotted with an offset of 1.0 between spectra.

Table 6. Spectroscopic parameters for both components of WASP 0639-32 obtained using equivalent width fitting, for cases where the surface gravity was free and where it was fixed at photometric values.

| Parameter | Primary Free | Secondary Free | Primary Fixed | Secondary Fixed |
|-----------|--------------|----------------|--------------|-----------------|
| $T_{\text{eff}}$ (K) | $6730 \pm 30$ | $5490 \pm 100$ | $5420 \pm 90$ | $5420 \pm 90$ |
| $\log g_s$ | $4.65 \pm 0.05$ | $3.97$ | $4.88 \pm 0.13$ | $4.61$ |
| $v_{\text{mic}}$ (km s$^{-1}$) | $1.61 \pm 0.01$ | $1.49 \pm 0.01$ | $2.67 \pm 0.18$ | $2.61 \pm 0.18$ |
| [Fe/H] | $-0.10 \pm 0.01$ | $-0.33 \pm 0.01$ | $-0.38 \pm 0.06$ | $-0.45 \pm 0.05$ |

EWs, $W_o$, to judge the choice of $\Theta_{\text{sp}}$. The overall log-likelihood function can be written as

$$
\ln L(W; \Theta_{\text{sp}}) = -\frac{1}{2} \sum_{i=1}^{N} \left( \frac{W_{o,i} - W_{s,i}(\Theta_{\text{sp}})}{\sigma_{W,i}} \right)^2 - \ln \left( \frac{2\pi}{\sigma_{W,i}} \right), \quad (8)
$$

where $W$ is a vector of length $N$ containing the fitted EWs, and $\sigma_{W,i}$ are the uncertainties associated with the measured EWs. Priorors were used to prevent the walkers exploring outside of the limits set by the model atmospheres, i.e., $4500 < T_{\text{eff}} < 7500$, $3.5 < \log g < 5.0$, $-1.0 < [\text{Fe/H}] < 1.0$, and $0 < v_{\text{mic}} < 100$. For both the primary and secondary star, 100 walkers and 500 steps were used. For both stars, the first 100 steps were removed to allow for adequate burn-in. Auto-correlation lengths are between 50 and 70 for the different runs, and acceptance fractions are 0.34–0.55.

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| $\log g_s$ | $4.65 \pm 0.05$ | $3.97$ | $4.88 \pm 0.13$ | $4.61$ |
| $v_{\text{mic}}$ (km s$^{-1}$) | $1.61 \pm 0.01$ | $1.49 \pm 0.01$ | $2.67 \pm 0.18$ | $2.61 \pm 0.18$ |
| [Fe/H] | $-0.10 \pm 0.01$ | $-0.33 \pm 0.01$ | $-0.38 \pm 0.06$ | $-0.45 \pm 0.05$ |

90 transiting planet host stars and showed that the discrepancy is temperature dependent. For a star with an effective temperature near 6700 K, the difference is typically 0.6 dex, which is consistent with what we have found for the primary of WASP 0639-32. In comparison, the difference between the two surface gravities for the secondary is smaller at 0.27 dex. This is larger than the trend of Mortier et al. (2013), but maybe caused by the small number of Fe II lines. Doyle (2015) notes that temperatures determined via the ionisation balance method are not affected by the discrepancy in log g but as our EW fitting includes microturbulence and metallicity, the different surface gravities may affect these parameters. We therefore ran the EW fitting procedure with log g fixed at 3.97 and 4.61, for the primary and secondary, respectively. The results are also shown in Table 6. By fixing the surface gravities of both stars, the effective temperature of primary has been reduced, while it has increased for the secondary. For both stars, the metallicity has been reduced. The temperature ratio, $T_2/T_1$, has increased from 0.816 ± 0.019 to 0.857 ± 0.016. Note, the uncertainties on the values in Table 6 do not include uncertainties from systematics such as those from using one specific line-list, solar abundance list or set of model atmospheres. Jofré et al. (2017) provide a summary of expected uncertainty contributions to the metallicity for systematic, such as these, with one on the largest contributions being the continuum placement with a systematic of 0.3 dex. As the uncertainty on the effective temperature of the primary star is very small and does not include potential systematics (as previously discussed), we take the uncertainty as ±0.5 K. This has been estimated from variations in the temperature when the segment width for the generated spectrum is altered slightly. This paper has used the
The equivalent width fitting gives two slightly different effective temperatures depending on whether we choose to use surface temperatures depending on whether we choose to use surface gravity and metallicity agree with the literature values. The microturbulence is within 2-σ uncertainty of the 15.9 and 84.1 percentiles. Overall the measured effective temperature, surface gravity and metallicity are shown in Table 8. The parameters themselves are taken as the median value from the distribution, with uncertainties calculated using the 15.9 and 84.1 percentiles. Overall the measured effective temperature, surface gravity and metallicity agree with the literature values. The microturbulence is within 2-σ of the literature value, which is itself a mean value obtained from multiple techniques.

5.4. Checking effective temperatures

The EW fitting gives two slightly different effective temperatures depending on whether we choose to use surface gravities.
from the spectroscopy, log $g_\ast$, or from the masses and radii, log $g_{\text{MR}}$. Therefore other methods of determining the effective temperature have been considered.

One of the spectra for WASP 0639-32, (taken at HJD = 2 457 042.59910) was taken during the flat region of the secondary eclipse when only the primary star would have been visible, therefore, it can treated in the same way as a single star spectrum. Figure 8 shows how multiple synthetic spectra, generated using different effective temperatures, compare to the $H_\alpha$ region of the eclipse spectrum. By looking at the wings of the $H_\alpha$ line, we found the best temperature was $6150 \pm 150$ K, which is lower than both estimates from the EW fitting. A surface gravity of 3.97 was used to generate the synthetic spectra, but as this method is insensitive to the surface gravity (Niemczura et al. 2014), the chosen value will not affect the best model spectrum. [Fe/H] was set to $-0.14$ ± 0.30, and $v_{\text{mac}} = 1.42$ ± 0.80 km s$^{-1}$. While the uncertainties are much larger than those obtain through the disentangling, the effective temperature has better agreement with the temperature from the disentangling with fixed surface gravity rather than the temperature from the $H_\alpha$ fitting.

We have also used empirical colour-effective temperature and colour-surface brightness relations to estimate the effective temperatures of the individual stars in the binary system, as described in Maxted & Hutcheon (2018). Photometry was taken from the AAVSO Photometric All-Sky Survey, APASS ($B, V, g', r', i'$), the Two Micron All Sky Survey, 2MASS ($JHK_s$), Tycho-2 Catalogue ($B_T$ and $V_T$), Deep Near-infrared Southern Sky Survey, DENIS, ($I, J, K$), and the Wide-Field Infrared Survey Explorer, WISE ($W_1$). See Table 9 for the observed values (Epchtein et al. 1997; Hög et al. 2000; Skrutskie et al. 2006; Henden et al. 2009; Wright et al. 2010).

To fit the observed photometry, the model uses the following parameters— $E(B-V)$, the apparent $g'$-band magnitudes for stars $i = 1, i = 2$ corrected for extinction; $T_{\text{eff}}$, the effective temperatures for each star; $E(B-V)$, the reddening to the system; $\sigma_{\text{ext}}$, the additional systematic error added in quadrature to each measurement to account for systematic errors. Although not used here, the model can work with triple systems, in which case there is a third $g'$-band and $T_{\text{eff}}$. The apparent magnitudes for each star are predicted through the colour-temperature relations of Boyajian et al. (2013). Following the example of Boyajian et al. (2013), the transformations of Bessell & Brett (1988) and Carpenter (2001) are used to transform between the photometric systems of Johnson and 2MASS. Cousins $I_c$ is used as an approximation to the DENIS Gunn $i'$ band and the 2MASS $K_s$ is used as an approximation to the DENIS $K$ band. Transformations between Johnson $B, V$, and Tycho $B_T$ and $V_T$ magnitudes are carried out by interpolating values in Table 3 of Bessell (2000). $V$-band extinction is assumed to be $3.1 \times E(B-V)$, extinction in the SDSS and 2MASS bands is calculated using $A_r = 2.770 \times E(B-V)$ from Fiorucci & Munari (2003), and the extinction coefficients to the $r'$ band come from Davynt et al. (2014). The average surface brightness ratio in the WASP band (0.429 ± 0.007) as an approximation to the $V$-band ratio, and allowed the external error $\sigma_{\text{ext}}$ account for differences in the two bands. This ratio is used as a constraint, by using the $V$-band surface brightness

![Fig. 8. Synthetic spectra of different temperatures plotted against the $H_\alpha$ wings of the primary star.](image-url)
Fig. 9. Difference between the observed magnitude and those obtained through the fitting of colour-temperature relations. Uncertainties on the y-axis include the external error, $\sigma_\text{ext}$, added in quadrature.

and B–K relations of Graczyk et al. (2017). The parameter-space was explored using a maximum likelihood approach, using an emcee (Foreman-Mackey et al. 2013) MCMC procedure, with 64 walkers, 1088 steps and 64 burn-in steps. All chains have visually inspected to ensure suitable mixing, and via a running mean check to ensure convergence. The median acceptance fraction for the walkers was 0.392. The total line-of-sight reddening from the dust maps of Schlafly & Finkbeiner (2011) is used to place a prior on $\Delta = E(B - V) - E(B - V)_\text{map}$:

$$P(\Delta) = \begin{cases} 1 & \Delta \leq 0 \\ \exp(-0.5(\Delta/0.034)^2) & \Delta > 0 \end{cases}$$

The 0.034 is a constant from Maxted et al. (2014), and is the offset between reddening maps and the values derived from photometry of a number of A-type stars.

The resulting temperatures were $T_1 = 6340 \pm 190$ K and $T_2 = 5320 \pm 170$ K with reddening of $E(B - V) = 0.067 \pm 0.036$ and $\sigma_\text{ext} = 0.043 \pm 0.017$. The parameter values are as the median of the posterior distributions, while the uncertainties are the standard deviations. Figure 9 shows the residuals from the fit.

Overall, by taking the weighted mean of the temperatures from each technique for each star (excluding spectroscopy with a free surface gravity), the effective temperatures are found to be $6330 \pm 50$ K and $5400 \pm 80$ K for the primary and secondary, respectively. Note that in these calculations, an additional 100 K has been added to the uncertainty in the temperature for the primary star obtained from the $H_\alpha$ fitting, due to the uncertainty in the continuum placement, and the uncertainty from EW fit with a fixed log $g$ has been taken as 50 K.

6. Absolute parameters

JKTABSDIM\(^3\) was used to calculate the absolute parameters for WASP0639-32, taking the relative radii and inclination from the detrended fit in Table 3, and the semi-amplitude velocities and eccentricity from Table 5. The results are presented in Table 10. log $g_{\text{MR}}$ is used to denote a surface gravity obtained directly from the mass and radius of the star. JKTABSDIM also calculates a distance estimate using the surface-brightness relations of Kervella et al. (2004). The 2MASS $K_s$ apparent magnitude has been used to estimate a distance, $d_K$, to WASP0639-32

\(^3\) http://www.astro.keele.ac.uk/jkt/codes/jktabsdim.html

and is shown in Table 10. The magnitude was converted to the Johnson system using equations from Bessell & Brett (1988) and Carpenter (2001), and the APASS V-band magnitude, giving $K = 9.267 \pm 0.097$ mag. Although other apparent magnitudes are available and give similar distances, for example, $a_d = 310 \pm 23$ pc and $d_d = 322 \pm 21$ pc for APASS BV, respectively, the $K$-band distance was chosen as it is the least sensitive to reddening. Also listed in this table, is the parallax from Gaia DR1 (Gaia Collaboration et al. 2016) and its associated Gaia distance, $d_{\text{Gaia}}$, for comparison. The distance from the $K$-band magnitude agrees well with $d_{\text{Gaia}}$.

The quoted uncertainties in the masses are 0.37% and 0.36% for $M_1$ and $M_2$, respectively, while the uncertainties in the radii are 1.27% and 1.12% for $R_1$ and $R_2$, respectively. For the masses, this is the same level of precision as was obtained for the evolved binary, AI Phe (Kirkby-Kent et al. 2016), meaning that like AI Phe, the parameters of WASP 0639-32 will be suitable for constraining the helium abundance in stellar evolutionary models. Figure 10\(^4\) shows the correlation between $M_1$ and $M_2$, calculated from $K_1$, $K_2$, and $e$ values explored during the MCMC in Sect. 4. The formulae recommended in Torres et al. (2010) were used for the calculation as they are as same as the formulae used in JKTABSDIM. The period and inclination were fixed for this calculation.

The projected rotational velocity, $v \sin i$, of both stars has been measured using the cross-correlation technique (Bouvier 2013). The spectroscopic parameters from the fixed log $g$ solution in Table 6, were used to generate synthetic spectra with $v \sin i$ values ranging from 2.0 to 9.0 km s\(^{-1}\). The spectra were generated using solar abundances from Asplund et al. (2009), MARCS.GES (Gustafsson et al. 2008) model atmospheres and the VALD line-list (Kupka et al. 2011). Each synthetic spectrum was cross-correlated with the NARVAL solar spectrum (Aurière 2003) within iSpec to set up a relation between $v \sin i$ and the full-width half-maximum (FWHM) of the cross-correlation peak. Each observed spectrum was cross-correlated with the same solar spectrum and the FWHM was obtained for both stars. After taking the mean and standard deviation of the resulting

\(^4\) Produced using corner.py (Foreman-Mackey 2016).

| Parameter | Primary | Secondary |
|-----------|---------|-----------|
| Period (days) | 11.658317(5) |          |
| $e$ | 0.0009(12) |          |
| $i$ (°) | 89.9943(67) |          |
| $a \sin i$ (R$_\odot$) | 26.964(31) |          |
| $a$ (R$_\odot$) | 26.964(31) |          |
| Mass ratio | 0.6785(16) |          |
| $M$ sin$^2 i$ (M$_\odot$) | 1.1544(43) 0.7833(28) |
| Mass (M$_\odot$) | 1.1544(43) 0.7833(28) |
| Radius (R$_\odot$) | 1.834(23) 0.7291(81) |
| log $g_{\text{MR}}$ | 3.974(11) 4.607(10) |
| $T_{\text{eff}}$ (K) | 6330(50) 5400(80) |
| log($L/L_\odot$) | 0.685(18) −0.392(28) |
| $d_K$ (pc) | 323(06) |
| Parallax (mas) | 3.10(23)(30) |
| $d_{\text{Gaia}}$ (pc) | 323(±23)±35 |
| $v \sin i$ (km s$^{-1}$) | 6.75(7) 4.7(1.1) |
| $\omega_{\text{synch}}$ (km s$^{-1}$) | 7.96(10) 3.16(4) |

Notes. Where two uncertainties are given, the first refers to the random error, and the second is the systematic.
Figs. 10. Density distribution showing the correlation between $M_1$ and $M_2$. A grey cross marks the measured values as shown in Table 10.

FWHMs, the established relation for each star was used to find the corresponding $v \sin i$, the results of which are shown in Table 10. Also included in Table 10 are the synchronous velocities $v_{\text{synch}}$ for both stars, which are calculated in JKTB-SDIM. Comparisons between $v \sin i$ and $v_{\text{synch}}$ show that both stars are almost rotating synchronously. For the primary star the difference is sufficient to say that is sub-synchronous, although there maybe systematic uncertainties that have not been taken into consideration, which may alter the outcome. One example would be the effect of microturbulence in the synthetic spectra. A possible explanation for the nature of the primary is that it is evolving on a timescale that is more rapid than the timescale for synchronisation.

7. Implication for stellar evolutionary models

7.1. The models

With uncertainties on the masses of less than 0.5% and one component evolved off of the main-sequence, WASP 0639-32 is a good binary system for testing stellar evolutionary models (Torres et al. 2010). The model tracks that we tested were produced using the Garching Stellar Evolution Code (GARSTEC, Weiss & Schlattl 2008) and include the initial helium abundance, $Y_i$, initial metallicity, $Z_i$, and mixing-length parameter, $\alpha_{\text{ml}}$, as free parameters.

The model grid covers six initial helium abundances, $Y_i = \{0.231, 0.251, 0.271, 0.291, 0.311, 0.331\}$ and five different mixing length parameters, $\alpha_{\text{ml}} = \{1.598, 1.698, 1.798, 1.898, 1.989\}$, yielding 30 combinations in total. For each pair of $Y_i$ and $\alpha_{\text{ml}}$, there are 15 values for $Z_i$, covering a range from 0.00307 to 0.0772 with equal spacing in log $Z$ space. For each point on an evolutionary track, $[\text{Fe/H}]_i$ is calculated using $[\text{Fe/H}]_i = \log_{10}(Z_i/X_i) - \log_{10}(0.02439)$, where 0.02439 is the solar value from Grevesse & Noels (1993). $Z_i$ and $X_i$ are the model metal fraction and model hydrogen fraction, respectively. The evolutionary tracks cover a mass range of 0.7–2.0 $M_\odot$ in steps of 0.02 $M_\odot$.

The microphysics used in these models is described in Serenelli et al. (2013) and Weiss & Schlattl (2008), but we provide a brief summary here. The convection is described by the standard mixing length theory of Kippenhahn & Weigert (1990), where the solar mixing length is $\alpha_{\text{ml,}\odot} = 1.801$ using the Grevesse & Noels (1993) solar composition. Due to the effects of diffusion, the initial solar composition is found to be $[\text{Fe/H}]_i = +0.06$. Molecular opacities from Ferguson et al. (2005) complement the OPAL opacities from Iglesias & Rogers (1996). Convective mixing is described in terms of a diffusive process, with the diffusion coefficient, $D_c$, being given by

$$D_c = \frac{1}{3} \alpha_{\text{ml}} H_p v_c,$$

where $v_c$ is the convective velocity and $\alpha_{\text{ml}} H_p$ is the local mixing length. Convective overshooting is described by the diffusive process (Freytag et al. 1996) given by

$$D = D_0 \exp\left(\frac{-2z}{f H_p}\right),$$

where $f$ is a free parameter dictating the extent of the overshoot, $D_0$ is the diffusion coefficient within the convective border and $z$ is the distance to the convective border. $H_p$ is a function of the pressure scale height ($H_p$) and the thickness of the convective core ($\Delta R_{\text{CZ}}$) given as

$$h_p = H_p \times \min\left[1, \left(\frac{\Delta R_{\text{CZ}}}{H_p}\right)^2\right].$$

This definition of $h_p$ ensures the overshooting region is limited geometrically to a portion of the convective region, in cases where the convective region is small e.g., where $\Delta R_{\text{CZ}} < H_p$ (Magic et al. 2010). If $\Delta R_{\text{CZ}} > H_p$, the geometric limit does not have an effect. In this case, our choice of $f = 0.020$ leads to an extension of the overshoot $\sim 0.25 H_p$.

7.2. Input data

For each star in WASP 0639-32, the vector of observed parameters is $d = (T, \rho, M, [\text{Fe/H}])$, where $T$ is the effective temperature, $\rho$ is the mean density, $M$ is the mass and [Fe/H] is the observed surface metal abundance. Using

$$\rho_n = \frac{3\pi}{G P^2(1 + Q_n) R_n^3},$$

the mean density is a quantity that can be calculated from the fractional radii $r_n = (R_n/a)$ obtained from the lightcurve solution and Kepler’s laws, meaning it is independent from the mass estimates obtained from the spectroscopic orbit. In this equation, $R$ is the radius for star $n = 1, 2$, $a$ is the semi-major axis.
of the orbit, \( P \) is the orbital period and \( G \) is Newton’s gravitational constant \( (\text{Maxted et al. 2015}). Q_b \) is a function of the mass ratio \( (q = M_2/M_1) \), where \( Q_1 = q \) and \( Q_2 = 1/q \). The densities of the two stars were found to be \( \rho_1 = 0.1873 \pm 0.0071 \rho_0 \) and \( \rho_2 = 2.023 \pm 0.073 \rho_0 \) for the primary and secondary respectively. By using \( M \) and \( \rho \) rather than \( M \) and \( R \), the parameters in the log-likelihood can be assumed to be uncorrelated. The masses used are those stated in Table 10 and the temperatures are the weighted means that calculated at the end of Sect. 5.4.

We use the metallicities from Table 6 using a fixed log \( g \), but assume uncertainties of \( \pm 0.1 \) dex to account take into consideration some of the uncertainties mentioned in Jofré et al. (2017). The grid of models described in Sect. 7.1 is quite coarse compared with the uncertainties on the masses and densities, and there is no interpolation between the models. As such, the uncertainties were increased to 0.0073 and 0.0050, for \( M_1 \) and \( M_2 \) respectively, to allow more evolutionary tracks to be explored. An uncertainty of 0.005 \( M_0 \), covers a range of \( \pm 2-\sigma \) from the centre of a mass bin.

7.3 Bayesian analysis

In order to explore this complex parameter space set out by the models, we use a Bayesian framework. The set of model parameters is given by \( m = (\tau, M, \alpha_{\text{mix}}, [\text{Fe}/\text{H}]) \), where \( \tau, M, \alpha_{\text{mix}} \), and \([\text{Fe}/\text{H}]\) are the mass, mixing length, and initial metal abundance respectively. The initial metal abundance differs from the observed surface abundance due to diffusion.

The probability distribution function \( p(m|d) \) is proportional to \( p(m) L(d|m) \). \( L(d|m) = \exp(-\chi^2/2) \) is the likelihood of observing the data \( d \) given the model \( m \), assuming the independent observed parameters and Gaussian uncertainties, where

\[
\chi^2 = \frac{(M - M_{\text{obs}})^2}{\sigma^2_M} + \frac{(\rho - \rho_{\text{obs}})^2}{\sigma^2_{\rho}} + \frac{(T - T_{\text{obs}})^2}{\sigma^2_T} + \frac{([\text{Fe}/\text{H}] - [\text{Fe}/\text{H}]_{\text{obs}})^2}{\sigma^2_{[\text{Fe}/\text{H}]}}. \tag{13}
\]

Observed parameters are denoted by the “obs” subscript and each \( \sigma \) represent the uncertainty of the parameter. \( p(m) \) is the product of the priors on each of the model parameters given by \( p(m) = p(\tau)p(M)p([\text{Fe}/\text{H}]) \). A flat prior is applied to the initial surface metallicity, although this usually has little effect as the surface metallicity is constrained by the observed value.

A very loose prior is used on the age aimed at keeping the star’s age within the age of the universe while allowing room for the models to explore slightly older ages in order provide a more realistic age estimate. The prior on the age is set to 0 < \( \tau \) < 17.5 Gyr.

Each \( Y \) value was considered independently and calculations were split into slices of age covering 300 Myr, e.g., 0–300 Myr, 300–600 Myr, etc. and the age is the weighted average over the bin. This is to allow comparisons between the two stars and to choose results where age and initial helium abundance were common to both star.

7.4. Model comparisons

Once the Bayesian analysis for each star and for each \( Y \) was complete, we used the best-fit age from the primary star’s results to define the age of the system, and then found the matching age to get the parameters for the secondary. The total chi-squared, \( \chi^2_{\text{tot}} \), is calculated by summing \( \chi^2_1 \) and \( \chi^2_2 \) from the matching age slice. The best-fit parameters from the corresponding age-slices are shown in Table 11 are for stars of the same age and with the same initial helium abundance. The primary star was chosen to define the age of the system as its age was constrained more by the observed parameters. Figure 11 demonstrates this, as \( \chi^2_1 \) (solid points) forms a much narrower minimum compared with \( \chi^2_2 \) (light grey open points) from the fit to the secondary star’s parameters. This is due to the evolved nature the primary star. Stars will spend only a small fraction of their life on the subgiant branch, which in itself is a strong constraint on the age of a subgiant star, and the tight observational constraints limit it further.

Individually, the hotter primary star strongly favours an initial helium abundance of \( Y_0 = 0.271 \). A higher helium abundance means a higher mean molecular weight, which in turn means a higher luminosity (Kippenhahn & Weigert 1990), and therefore temperature, when \( M \) and \( \rho \) are fixed. However, with small uncertainties on the temperatures providing such tight constraints, it is the parameters with weaker constraints, such as \( \alpha_{\text{mix}} \) and \([\text{Fe}/\text{H}]\) that have to adapt. From the solutions for the secondary star, the smallest \( \chi^2 \) is seen for \( Y_0 = 0.231 \), which is lower than that of the primary component. This solution sacrifices the temperature in order to improve the fit to the mass and radius. This value is in fact less than the primordial helium abundance at the time of the big-bang nucleosynthesis, \( Y_{\text{BBN}} = 0.2485 \) (Steigman 2010). If this solution is excluded for this reason, \( Y_0 = 0.271 \) is the next best solution for the secondary star. This matches the value found for the primary.

After summing the \( \chi^2 \) from the individual stars, the best helium abundance for the system as a whole is \( 0.271 \), although it could be argued that a range between 0.231–0.271 could provide reasonable solutions, using \( \Delta\chi^2_{\text{tot}} = 1 \) as an estimate of a 68% confidence interval (Press et al. 1992). We refrain from quoting a true confidence region due the discontinuous nature of the helium abundances tested, and because there is no interpolation between the individual tracks. Better solutions could exist between the models. Interpolation has not been employed in these models due to the large number of parameters which would need to be considered. If the same estimate of a confidence interval for \( \chi^2_1 \) is used then the range of helium abundances is \( Y_1 = 0.251–0.291 \), however \( Y_1 = 0.291 \) is disfavoured by the secondary due to a high temperature and high metallicity.

The lowest \( \chi^2_{\text{tot}} \) suggests the age of WASP0639-32 is 4.22 Gyr. Taking the uncertainty of \( Y \) into consideration, as \( \pm 1 \) in \( \chi^2_1 \), the overall age of the system is \( 4.2^{+0.8}_{-0.5} \) Gyr. The \( Y_1 = 0.231 \) solution was excluded as it is less than \( Y_{\text{BBN}} \). The age estimate does not consider the potential effects of different overshooting efficiencies, \( f \), or different parameterisations of the overshooting. The age of the primary star may be particularly sensitive to the overshooting, as its mass places the star on the boundary for developing convective cores (1.1–1.2 \( M_\odot \); Lebreton et al. 2014). A full investigation into the overshooting parameter is not possible with our current set of models, but there are a few simple tests to show how our choice of overshooting may affect the age estimate. For the tests, we consider a 1.16 \( M_\odot \) star, with \([\text{Fe}/\text{H}] = -0.3 \) and \( \alpha_{\text{mix}} = 1.70 \). These are the best tracks for the \( Y = 0.271 \) solution for the primary star. Figure 12 shows the tracks for three different overshoot parameterisations. The track marked “standard” is the same parameterisation that is described in Sect. 7.1 and that has been used throughout the modelling so far, i.e., for a 1.16 \( M_\odot \) star where the geometric cut plays a role and the overshooting is partially suppressed. We also present the
the mixing lengths compared to those by Magic et al. (2015), but
mixing-lengths, with the more evolved primary star requiring a
overshooting. The effects of mass uncertainties on the overshoot-
the observed masses would restrict the possible values for the
case where there is no overshooting and the case where the geo-
metric cut does not play a role. The removal of the geometric cut
means the overshooting is overestimated for the star, and
provides an upper limit that is approximately equal to 0.3 \( \Delta R_g \).

Figure 12 shows that removing the geometric cut results in a
track that is a poor choice for the primary star. It also shows
that reducing the level of overshooting will have little effect on
the resulting age of the system. The same trends are seen for
tracks with different helium abundances. Although, only one
mass-track has been considered here, the small uncertainties on
the observed masses would restrict the possible values for the
overshooting. The effects of mass uncertainties on the overshoot-
ning parameter has been discussed in Valle et al. (2017). It is worth
noting that any uncertainty in the overshooting used for the pri-
mary star could affect the parameters that are presented for the
secondary (as they are matched to the age of the primary star)
and therefore \( \chi^2_{\text{tot}} \). However, as discussed above, the effect should
be small.

For the lowest \( \chi^2_{\text{tot}} \) solution, the two stars use slightly different
mixing-lengths, with the more evolved primary star requiring a
value that is less than the solar value. The smaller secondary
component prefers a mixing length that is about solar. This is
consistent with the results from 3D radiative hydrodynamic mod-
els of Magic et al. (2015). There is some degeneracy between \( Y \)
and \( \alpha_{\text{ml}} \) as they both affect the temperature, but the extent of the
degeneracy is not fully understood. The degeneracy would need
to be explored in more detail, for a strong claim to be made about
the mixing lengths compared to those by Magic et al. (2015), but
this is beyond the capabilities of our current grid of models. One
point to note about this work is the different solar compositions
used between the spectroscopy (Asplund et al. 2009) and these
models (Grevesse & Noels 1993). This may introduce some
discrepancies in the modelling of the temperatures, although by
using a temperature that is obtained as a weighted mean from
multiple methods, this discrepancy should be reduced. As the
mixing length parameter has been included as a separate free
parameter for both stars, it is possible that these free parameters
are masking discrepancies between the models and observations,
i.e., by picking a higher/lower mixing length it is possible to find
a better match to the temperatures.

Figures 13 and 14 show the best fit evolutionary tracks for
the three best values of initial helium abundance, for each star.
The points are plotted with the temperatures and luminosities
from Table 10.

8. Discussion
The trends in \( \chi^2 \) in Table 11 are quite clear for each star, making
it possible to identify the preferred values for \( Y \). However, with
the priors on the mass being looser than the observed value,
there is a chance the fit to the mass has been sacrificed in order
to provide a better match to other parameters such as the tem-
peratures. A finer grid of models, particularly in regards to the mass
and density parameters, would allow the tighter observation
constraints to be used on these parameters and would enable
the helium abundance to be constrained further by making the best \( \chi^2 \) value more distinct. As a test, the tracks were also fitted using the measured uncertainties on the masses and densities, and this allowed \( Y_1 \) to be constrained to the \( Y_1 = 0.271 \) solution. However, the solutions presented very jagged fits in Table 11 for the primary star. In number. An alternative to being explored independently of the initial helium abundance. The lowest \( \chi^2 \) values occurred for mixing length that were solar \( \alpha = 1.78 \) and \( \Delta Y = 0.0 \), and yielding \( Y_i = 0.257 \) with a range of \( \pm 0.020 \). The two modelling methods use different chemical abundances meaning direct comparison should not be made, but it worth noting that this value matches the two initial helium abundances meaning direct comparison should not be made, but it worth noting that this value matches the two initial helium abundances with the lowest \( \chi^2 \) in Table 11. The spectroscopic parameters obtained using a free log \( g \) were also tested using the method of Kirkby-Kent et al. (2016) to see if the evolution of the system differs. With the primary star set with a temperature of 6730 K, a good fit could not be found. The best solutions were suggesting a mixing length larger than 2.32, and a helium abundance larger than 0.31. Even these solutions were very poor fits, with the models unable to match the densities and temperatures of both stars. This highlights the need to ensure accurate and consistent results for parameters that are used in stellar evolutionary modelling. It also highlights that caution should be used if the only source for a surface gravity is a spectroscopic value. Mortier et al. (2013) shows that this issue is most prevalent at temperatures that deviates significantly from the Sun, but as yet, an exact cause has not been found. Porto de Mello et al. (2008) looked at the atmospheric parameters for \( \alpha \) Centauri A and B. They noted a disagreement between photometric and spectroscopic effective temperature scales for \( \alpha \) Centauri A, with the spectroscopically derived temperature being higher than those obtained photometrically. This is similar to what has been found for the primary star of WASP 0639-32. Porto de Mello et al. (2008) suggest non-local thermodynamic equilibrium (NLTE) effects as a cause for the offset. As no NLTE effects are considered in the analysis in this paper, it is possible that this is the cause of the disagreement between the temperatures obtained with the fixed and free

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**Table 11: Best-fit evolutionary models for the primary and matching age model for the secondary, using different initial helium abundances \( Y_i \).**

| Parameter          | Unit | 0.231 | 0.251 | 0.271 | 0.291 | 0.311 | 0.331 | Observed Value | Error |
|--------------------|------|-------|-------|-------|-------|-------|-------|----------------|-------|
| \( T_{\text{best}} \) (Gyr) |       | 5.31  | 5.05  | 4.22  | 4.12  | 3.55  | 2.98  | –              | –     |
| \( T_i \) (K)       |       | 6317  | 6297  | 6330  | 6340  | 6380  | 6356  | 6330           | 50    |
| \([\text{Fe/H}]_{\text{ref}}\) – | –     | –0.47 | –0.35 | –0.35 | –0.23 | –0.23 | –0.22 | –0.33          | 0.10  |
| \( M_1 \) (\( M_\odot \)) |       | 1.1539| 1.1552| 1.1568| 1.1568| 1.1495| 1.1578| 1.1544         | 0.0073|
| \( R_1 \) (\( R_\odot \)) |       | 1.8324| 1.8281| 1.8331| 1.8431| 1.8053| 1.7897| 1.8345         | 0.023 |
| \( \mu_1 \) (\( \mu_\odot \)) |       | 0.1866| 0.1882| 0.1847| 0.1847| 0.1806| 0.2010| 0.1873         | 0.0071|
| \( \log g_1 \) – |       | 3.97  | 3.98  | 3.98  | 3.97  | 3.96  | 4.00  | 3.974          | 0.0111|
| \( \alpha_{\text{rad}} \) – | –     | 2.043 | 2.048 | 1.704 | 1.921 | 1.639 | 1.501 | –              | –     |
| \( \chi_i \) – | –     | 1.96  | 0.53  | 0.10  | 1.37  | 3.31  | 5.39  | –              | –     |
| \( T_2 \) (K)       |       | 5363  | 5397  | 5410  | 5480  | 5487  | 5520  | 5400           | 80    |
| \([\text{Fe/H}]_{\text{ref}} \) – | –     | –0.53 | –0.45 | –0.35 | –0.26 | –0.16 | –0.13 | –0.45          | 0.11  |
| \( M_2 \) (\( M_\odot \)) |       | 0.7784| 0.7746| 0.7766| 0.7874| 0.7995| 0.7975| 0.7833         | 0.0050|
| \( R_2 \) (\( R_\odot \)) |       | 0.7245| 0.7251| 0.7263| 0.7317| 0.7322| 0.7339| 0.7291         | 0.0081|
| \( \mu_2 \) (\( \mu_\odot \)) |       | 2.041 | 2.027 | 2.021 | 1.982 | 1.979 | 1.966 | 2.023          | 0.073 |
| \( \log g_2 \) – |       | 4.61  | 4.61  | 4.61  | 4.60  | 4.60  | 4.60  | 4.607          | 0.010 |
| \( \alpha_{\text{rad}} \) – | –     | 1.810 | 1.864 | 1.880 | 1.962 | 1.973 | 1.754 | –              | –     |
| \( \chi_i \) – | –     | 1.64  | 2.83  | 2.53  | 5.07  | 8.89  | 11.67 | –              | –     |
| \( \chi_{\text{tot}} \) – | –     | 3.60  | 3.36  | 2.63  | 6.44  | 12.20 | 17.06 | –              | –     |

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**Fig. 14.** Evolutionary tracks in the temperature-density plane for the primary (black) and secondary (grey) components for different helium abundances. Dashed, \( Y = 0.291 \); solid, \( Y = 0.271 \); dotted, \( Y = 0.251 \); dot-dashed, \( Y = 0.231 \). All tracks are plotted from an age of 35 Myr. Tracks closest the observed mass of each star are plotted.
surface gravities. With further data releases to come from Gaia, it will be possible to obtain model independent effective temperature measurements using the angular diameter and integrated fluxes of the stars.

9. Conclusion

We have determined the masses and radii of the stars in a previously unstudied binary system, WASP-0639-32, to the precision of 0.37% and 0.36% for \( M_1 \) and \( M_2 \) respectively and 1.27% and 1.12% for the \( R_1 \) and \( R_2 \) respectively. Taking the weighted mean of the temperature from each technique for each star (excluding spectroscopy with a free surface gravity), the effective temperatures are found to be 6330 \( \pm \) 50 K and 5400 \( \pm \) 80 K for the primary and secondary, respectively.

With precise measurements of its fundamental parameters, this detached eclipsing binary system is a good choice for calibrating both main-sequence evolutionary models using the secondary star, but also subgiant models using the primary star. We have shown that it is possible to constrain the helium abundance of the system to a range 0.251–0.271, with the possibility of constraining it further with a finer grid of models.

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References

Asplund, M., Grevesse, N., Sauval, A. J., & Scott, P. 2009, ARA&A, 47, 481
Aurière, M. 2003, EAS Pub. Ser., 9, 105
Bailer, P., Modigliani, A., Bonisman, O., et al. 2000, The Messenger, 101, 31
Bessell, M. S. 2000, PASP, 112, 961
Bessell, M. S., & Brett, J. M. 1988, PASP, 100, 1134
Blanco-Cuaresma, S., Soubiran, C., Heiter, U., & Jofré, P. 2014a, A&A, 569, A111
Blanco-Cuaresma, S., Soubiran, C., Jofré, P., & Heiter, U. 2014b, A&A, 566, A98
Bouvier, J. 2013, EAS Pub. Ser., 62, 143
Boyajian, T. S., von Braun, K., van Belle, G., et al. 2013, ApJ, 771, 40
Brogard, K., Vandenberg, D. A., Bruntt, H., et al. 2012, A&A, 541, A106
Burgers, J. M. 1969, Flow Equations for Composite Gases (New York: Academic Press)