The Use of the Revised Trauma Score as an Entry Criterion in Traumatic Hemorrhagic Shock Studies: Data from the DCLHb Clinical Trials

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Abstract
Introduction: The Revised Trauma Score (RTS) has been proposed as an entry criterion to identify patients with mid-range survival probability for traumatic hemorrhagic shock studies.

Hypothesis/Problem: Determination of which of four RTS strata (1-3.99, 2-4.99, 1-4.99, and 2-5.99) identifies patients with predicted and actual mortality rates near 50% for use as an entry criterion in traumatic hemorrhagic shock clinical trials.

Methods: Existing database analysis in which demographic and injury severity data from two prior international Diaspirin Cross-Linked Hemoglobin (DCLHb) clinical trials were used to identify an RTS range that could be an optimal entry criterion in order to find the population of trauma patients with mid-range predicted and actual mortality rates.

Results: Of 208 study patients, the mean age was 37 years, 65% sustained blunt trauma, 49% received DCLHb, and 57% came from the European Union study arm. The mean values were: ISS, 31 (SD 18); RTS, 5.6 (SD 1.8); and Glasgow Coma Scale (GCS), 10.4 (SD 4.8). The mean TRISS-predicted mortality was 34% and the actual 28-day mortality was 35%. The initially proposed 1-3.99 RTS range (n = 41) had the highest predicted (79%) and actual (71%) mortality rates. The 2-5.99 RTS range (n = 79) had a 62% predicted and 53% actual mortality, and included 76% blunt trauma patients. Removal of GCS <5 patients from this RTS 2-5.99 subgroup caused a 48% further reduction in eligible patients, leaving 41 patients (20% of 208 total patients), 66% of whom sustained a blunt trauma injury. This subgroup had 54% predicted and 49% actual mortality rates. Receiver operator curve (ROC) analysis found the GCS to be as predictive of mortality as the RTS, both in the total patient population and in the RTS 2-5.99 subgroup.

Conclusion: The use of an RTS 2-5.99 inclusion criterion range identifies a traumatic hemorrhagic shock patient subgroup with predicted and actual mortality that approach the desired 50% rate. The exclusion of GCS <5 from this RTS 2-5.99 subgroup patients yields a smaller, more uniform patient subgroup whose mortality is more likely related to hemorrhagic shock than traumatic brain injury. Future studies should examine whether the RTS or other physiologic criteria such as the GCS score are most useful as traumatic hemorrhagic shock study entry criteria.

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the study fluid as compared to use of standard therapies. Optimally, the ability to detect the efficacy of a new therapy requires study subjects with predicted mid-range mortality and Trauma and Injury Severity Score (TRISS) survival predictions of 40 to 60%. 

Both Diaspirin Cross-Linked Hemoglobin (DCLHb, Baxter Healthcare Corporation, Chicago, Illinois USA) and PolyHeme (Northfield Laboratories Inc., Evanston, Illinois USA) have been studied in the prehospital setting utilizing physiologic criteria indicative of hypoperfusion as the entry criteria. A previously proposed prehospital traumatic hemorrhagic shock study of HBOC-201 (Biopure Corporation, Cambridge, Massachusetts USA) suggested the use of an RTS stratum of 1-3.99 as a proposed entry criterion. Use of RTS as a study entry criterion allows for ease of prehospital calculation using three clinical variables that can be rapidly assessed: Glasgow Coma Scale (GCS), systolic blood pressure (SBP), and respiratory rate (RR). RTS has been examined with regard to predicting mortality in trauma patients, injury severity, outcome of polytraumatized patients, and length of hospital stay. However, no studies have examined RTS as a study entry criterion, nor has any traumatic hemorrhagic shock clinical trial specifically utilized RTS as an entry criterion.

The purpose of this study was to compare the proposed RTS 1-3.99 range with other RTS ranges in order to determine which

| Table 1. Patient Demographics and Clinical Variables by Study Site |
|---------------------------------------------------------------|
| Abbreviations: DCLHb, Diaspirin cross-linked hemoglobin; GCS, Glasgow Coma Scale; ISS, Injury Severity Score; MOI, Mechanism of Injury; ns, not significant; NS, Normal Saline; RR, Respiratory Rate; RTS, Revised Trauma Score; SBP, Systolic Blood Pressure; TRISS, Trauma and Injury Severity Score. |
| *P = ns for US vs. EU. |

|                          | US     | EU     | Total  | P   |
|--------------------------|--------|--------|--------|-----|
| n (%)                    | 90 (43.3) | 118 (56.7) | 208 (100) | —   |
| Age, mean (SD)           | 39.0 (19.7) | 35.4 (13.9) | 37.0 (13.9) | .12 |
| Gender                   |        |        |        |     |
| Male, n (%)              | 69 (76.7) | 81 (68.6) | 150 (72.1) | .20 |
| Female, n (%)            | 21 (23.3) | 37 (31.4) | 58 (27.9)  |     |
| Treatment                |        |        |        |     |
| DCLHb, n (%)             | 47 (52.2) | 54 (45.8) | 101 (48.6) | ns  |
| NS, n (%)                | 43 (47.8) | 64 (54.2) | 107 (51.4) |     |
| MOI                      |        |        |        |     |
| Blunt, n (%)             | 53 (58.9) | 82 (69.5) | 135 (64.9) | .11 |
| Penetrating, n (%)       | 37 (41.1) | 36 (30.5) | 73 (35.1)  |     |
| ISS, mean (SD)           | 30.8 (16.7) | 30.4 (18.2) | 30.6 (17.5) | ns  |
| RTS, mean (SD)           | 5.43 (2.0) | 5.77 (1.6) | 5.62 (1.8)  | .17 |
| GCS Score                |        |        |        |     |
| mean (SD)                | 9.83 (5.3) | 10.9 (4.4) | 10.4 (4.8)  | .11 |
| median                   | 12      | 12      | 12      |     |
| mode                     | 15      | 15      | 15      |     |
| Entry SBP, mean (SD)     | 79.3 (16.1) | 73.1 (21.0) | 75.8 (19.2) | .02 |
| Entry RR, mean (SD)      | 22.1 (7.4) | 19.7 (7.0) | 20.0 (7.1)  | .15 |
| TRISS-Predicted Survival Rate, %, mean (SD) | 62.0 (36.8) | 69.6 (32.6) | 66.2 (34.7) | .13 |
| Mortality                |        |        |        |     |
| Predicted                | 38.0%   | 30.4%   | 33.8%   |     |
| Actual                   | 30.0% (27/90) | 39.0% (46/118) | 35.1% (73/208) |     |

P = ns for US vs. EU.
RTS range would identify a population of blunt and penetrating trauma patients with predicted and actually mortality rates near 50%. Additionally, receiver operator curve (ROC) calculations compared the composite RTS score and its component variables (GCS, SBP, RR) to understand which of these four variables best predicted 28-day mortality.

Methods

Data for this analysis of RTS stratifications comes from the paired, multi-center, randomized, single-blinded, normal saline-controlled, phase III clinical efficacy and safety studies of DCLHb in severe traumatic hemorrhagic shock. The US study involved 98 patients enrolled in the efficacy trial in 17 US trauma centers from February 1997 through January 1998; the European Union (EU) study enrolled 121 patients in four Belgian, 17 French, and 11 German trauma centers from July 1997 through May 1998.\(^\text{14,26}\) Inclusion criteria required that patients have hemorrhage and proven hypoperfusion (SBP < 90 mm Hg and HR > 120 beats/min, SBP < 90 mm Hg and HR < 60 beats/minute, or base deficit > 15 mEq/L). Patients excluded from the studies were those with traumatic brain injury, patients with imminent death, patients whose injury occurred more than four hours prior to infusion, minors, and pregnant women. The patients from the two traumatic hemorrhagic shock clinical trials were utilized for this analysis because they most

|                | DCLHb | NS | Total | \(P\) |
|----------------|-------|----|-------|------|
| n (%)          | 101 (48.6) | 107 (51.4) | 208 (100) | —    |
| Age, mean (SD) | 36.7 (17.8) | 37.19 (15.8) | 37.0 (13.8) | .84  |
| Gender         |       |     |       |      |
| Male, n (%)    | 75 (74.3) | 75 (70.1) | 150 (72.1) | .50  |
| Female, n (%)  | 26 (25.7) | 32 (29.9) | 58 (27.9)  |      |
| Study Site     |       |     |       |      |
| US, n (%)      | 47 (46.5%) | 43 (40.2) | 90 (43.3)  | .36  |
| EU, n (%)      | 54 (53.5%) | 64 (59.8) | 118 (56.7) |      |
| MOI            |       |     |       |      |
| Blunt, n (%)   | 65 (58.9) | 70 (69.5) | 135 (64.9) | .87  |
| Penetrating, n (%) | 36 (41.1) | 37 (30.5) | 73 (35.1)  |      |
| ISS, mean (SD) | 31.6 (19.1) | 29.6 (15.9) | 30.6 (17.5) | .42  |
| RTS, mean (SD) | 5.52 (1.84) | 5.72 (1.72) | 5.62 (1.80) | .41  |
| GCS Score      |       |     |       |      |
| mean (SD)      | 10.3 (4.9) | 10.6 (4.8) | 10.4 (4.8) | .65  |
| Median         | 12     | 12 | 12    |      |
| Mode           | 15     | 15 | 15    |      |
| Entry SBP, mean (SD) | 77.0 (17.1) | 74.7 (21.0) | 75.8 (19.2) | .42  |
| Entry RR, mean (SD) | 20.3 (7.2) | 19.9 (7.1) | 20.0 (7.1) | .71  |
| TRISS-Predicted Survival Rate, %, mean (SD) | 63.3 (35.8) | 69.0 (33.4) | 66.2 (34.7) | .25  |
| Mortality      |       |     |       |      |
| Predicted      | 36.7%  | 31.0% | 33.8%  |      |
| Actual         | 42.6% (43/101) | 28.0% (30/107) | 35.1% (73/208) | .028 |

\(^{a}\)\(P = .028\) for DCLHb vs. NS.

Table 2. Patient Demographics and Clinical Variables by Treatment Group

Abbreviations: DCLHb, Diaspirin cross-linked hemoglobin; GCS, Glasgow Coma Scale; ISS, Injury Severity Score; MOI, Mechanism of Injury; NS, Normal Saline; RR, Respiratory Rate; RTS, Revised Trauma Score; SBP, Systolic Blood Pressure; TRISS, Trauma and Injury Severity Score.
accurately represented the population of patients that may be studied in future resuscitation trials using hemoglobin-based solutions.

The data for the current retrospective database analysis study came from the original datasets provided by Baxter Healthcare from the US and EU studies. The RTS was calculated with entry GCS, SBP, and RR, utilizing the formula RTS = 0.9368(GCS) + 0.7326(SBP) + 0.2908(RR), using codified variables in the equation (www.trauma.org/archive/scores/rts.html). All SBP measurements are in millimeters of mercury (mm Hg) and all RR measurements are given in breaths per minute (breaths/min). Final patient survival status (lived vs. died) was based on all-cause 28-day mortality. The analyses included three patients for whom 28-day mortality status was unknown, with the assumption that they were alive at 28 days.

Statistical analysis of the RTS data included mean and standard deviation for age, Injury Severity Score (ISS), RTS, GCS, entry SBP and RR, and the TRISS-predicted survival. Analysis of variance, Chi-square, and two sample mean and proportion tests of significance were used at the P<.05 significance level. Receiver operator curve (ROC) analysis compared the capability of the composite RTS and its component variables in predicting 28-day mortality, using IBM SPSS Statistics v20.0 (IBM Corporation, Armonk, New York USA) and Epi Info StatCalc v3.5.1 (Centers for Disease Control and Prevention, Atlanta, Georgia USA).

Initially, the RTS 1-3.99 stratum proposed in the HBOC-201 RESUS clinical trial protocol was analyzed, with subsequent analyses of strata that were higher in value (RTS 2-4.99), wider in value (1-4.99), and both higher and wider in RTS value (2-5.99). Because the RTS 2-5.99 stratum had optimal TRISS-predicted and actual mortality rates, it was further analyzed in an effort to better understand if this range would best identify patients that would make the at the time pending HBOC-201 RESUS study and other traumatic hemorrhagic shock clinical trials feasible.

The protocols used in the US and EU clinical trials were approved by the Institutional Review Board (IRB) of each participating institution prior to the enrollment of any subjects. Trials were conducted in compliance with all regulations for good clinical trials and practice. The US study was conducted under federal regulations governing emergency research with an exception to informed consent. The current analysis of the data was conducted with IRB approval.

Results
Included in the study were 208 patients (95%) from the two DCLHb clinical trials for whom a valid RTS score was available. Mortality and treatment group did not differ in the 11 patients for whom the RTS was not available.

The mean age of study patients was 37 (SD = 14) years; 65% of the patients sustained a blunt injury, 49% received DCLHb resuscitation, and 57% were studied in the European Union (Table 1). The mean ISS was 31 (SD = 18), RTS was 5.62 (SD = 1.8), GCS was 10.4 (SD = 4.8), SBP was 76 (SD = 19), and RR was 20 (SD = 7). The TRISS-predicted survival rate for the entire study group was 66% (SD = 35%). This gave a TRISS-predicted mortality rate of 34% which did not differ from the actual mortality rate of 35%. Aside from a slightly higher entry SBP in the US patients as compared to EU patients (79 vs. 73, P<.024), there were no significant differences based on study site.

The only difference in demographics or clinical variables based on treatment with DCLHb was a 52% higher actual mortality rate in patients treated with DCLHb as compared to NS (P = 0.03) (Table 2). Of the 135 combined dataset patients (65%) who survived to 28 days, differences were observed based on patient outcome in patient age, mechanism of injury (MOI), ISS, RTS, GCS, DCLHb treatment, and TRISS–predicted survival (Table 3). Non-survivors had a lower RTS (4.61 vs. 6.17), a lower mean GCS score (7.96 vs. 11.8), and a lower TRISS-predicted survival rate of 40% vs. 79% (P<.001). The TRISS-predicted survival prediction distribution data demonstrated a bimodal patient distribution, with 106 patients (53%) in the 80-100% TRISS-predicted survival group and 37 (19%) in the 0-20% TRISS-predicted survival group, totaling 72% of the total patient population (lightest bars, Figure 1). Only 11 patients (5.5%) fell into the mid-range mortality range of 40-60% predicted survival. Actual and TRISS-predicted survival within all of the TRISS-predicted survival strata were comparable except in the 0-20% predicted survival stratum, in which actual survival was higher than the TRISS-predicted survival rate.

Patients were stratified into different RTS ranges in order to determine which subgroup provided the greatest number of traumatic hemorrhagic shock patients with predicted and actual mortality rates near 50%. The stratifications included the initial RTS range of 1-3.99, a higher range (2-4.99), a wider range (1-4.99), and a higher and wider range (2-5.99). The RTS 2-5.99 subgroup provided the largest number of patients (n = 79, 38%), the greatest proportion of penetrating injury patients (24%), the highest mean RTS (3.99, SD = 1.1), the highest mean GCS (5.85, SD = 3.5), the highest TRISS-predicted survival rate (38%, SD = 30%), and the actual mortality rate (53%) closest to the desired 50% mortality rate (Table 4). Predicted and actual mortality rates did not differ within any of the RTS strata subgroups.

Unlike the overall patient population, the RTS 2-5.99 subgroup did not show a bimodal TRISS-predicted survival curve (white bars, Figure 2). In this subgroup, 39% of the patients were in the 0-20% TRISS-predicted survival stratum and 71% were evenly distributed through the other strata. As such, the percentage of patients in the 20%-100% TRISS-predicted survival stratum significantly decreased upon isolation of the
RTS 2–5.99 patients as compared to the total patient population (61% vs. 81%, \( P < .001 \)). In the RTS 2–5.99 subgroup, 47% of patients had a TRISS-predicted mortality between 20% and 80%, while in the total patient population only 28% had a predicted survival within this range (\( P < .004 \)).

Due to its favorable mortality characteristics, further analysis of the RTS 2–5.99 subgroup was performed. United States patients in this RTS 2–5.99 subgroup had a lower mortality rate (39% vs. 65%, \( P < .02 \)) (Table 5). Patients in the RTS 2–5.99 subgroup who survived their trauma were more likely to have sustained penetrating trauma (43% vs. 7.1%), had a lower mean ISS (28 vs. 45), a higher RTS (4.33 vs. 3.68), and a higher TRISS-predicted mean survival rate (55% vs. 21%) as compared to non-survivors in this subgroup (\( P < .01 \)). In the RTS 2–5.99 subgroup, there was no difference in age, gender distribution, DCLHb treatment, GCS, SBP, or RR between patients who survived and expired by 28 days.

When comparing RTS 2–5.99 patients based on mechanism of injury, penetrating trauma patients had a lower mean ISS (21 vs. 42), a higher mean RTS (4.67 vs. 3.77), a higher mean GCS (7.89 vs. 5.20), and a higher mean TRISS-predicted survival rate (62% vs. 29%) as compared to blunt trauma patients in this RTS stratum (\( P < .003 \)) (Table 6). Blunt injury patients in this RTS stratum had a 1.9x higher predicted mortality rate (71% vs. 38%), and a 4.1x higher actual mortality rate (65% vs. 16%) as compared to penetrating trauma patients (\( P < .001 \)). Predicted and actual mortality in this RTS 2–5.99 stratum did not differ in either penetrating or blunt patient subgroups.

The RTS 2–5.99 subgroup patients had a higher RTS (3.99 vs. 2.77), a higher GCS (5.85 vs. 3.49), a higher RR (17 vs. 12),

| n (%) | Survivors | Nonsurvivors | Total | \( P \) |
|-------|-----------|--------------|-------|--------|
| Age, mean (SD) | 35.0 (13.2) | 40.6 (21.5) | 37.0 (16.8) | .023 |

**Table 3. Patient Demographics and Clinical Variables by Outcome**

Abbreviations: DCLHb, Diaspirin cross-linked hemoglobin; GCS, Glasgow Coma Scale; ISS, Injury Severity Score; MOI, Mechanism of Injury; ns, not significant; NS, Normal Saline; RR, Respiratory Rate; RTS, Revised Trauma Score; SBP, Systolic Blood Pressure; TRISS, Trauma and Injury Severity Score.
and a higher mean TRISS-predicted survival rate (38% vs. 21%) as compared to the RTS 1-3.99 subgroup patients \((P < .02)\). Among survivors, the RTS 2-5.99 subgroup patients had a higher RTS (4.34 vs. 3.05), a higher GCS (6.46 vs. 3.75), and a higher TRISS-predicted survival rate (35% vs. 32%; \(P < .03\)). Among non-survivors, the RTS 2-5.99 subgroup patients had a higher RTS (3.68 vs. 2.65), and a higher GCS (5.31 vs. 3.38) as compared to the RTS 1-3.99 subgroup \((P < .002)\).

Because GCS <5 patients have been shown to confound the ability of traumatic hemorrhagic shock clinical trials to determine therapy efficacy, further analyses of the RTS stratum with the exclusion of GCS <5 patients was conducted. The RTS 2-5.99 subgroup distribution with the exclusion of GCS <5 patients showed 18% of patients in the 40%-60% TRISS-predicted survival subgroup, displaying a uniform distribution across all TRISS-predicted survival strata (Figure 3). Without the GCS <5 patients, there was no significant increase in the distribution of patients with TRISS-predicted survival rates above 20% as compared to the RTS 2-5.99 subgroup with GCS <5 patients included (72% vs. 61%, \(P = ns\)).

Table 4. Clinical Data Based on RTS Strata from the Two DCLHb Traumatic Hemorrhagic Shock Clinical Trials

|                     | Total       | 1-3.99     | 2-4.99     | 1-4.99     | 2-5.99     |
|---------------------|-------------|------------|------------|------------|------------|
| n (%)               | 208 (100)   | 41 (19.7)  | 59 (28.4)  | 64 (30.8)  | 79 (38.0)  |
| Age, mean (SD)      | 37.0 (16.8) | 38.0 (20.0)| 39.6 (19.8)| 39.7 (19.6)| 39.3 (19.2)|
| MOI                 |             |            |            |            |            |
| Blunt, n (%)        | 135 (64.9)  | 37 (90.2)  | 48 (81.4)  | 53 (82.8)  | 60 (75.9)  |
| Penetrating, n (%)  | 73 (35.1)   | 4 (9.8)    | 11 (18.6)  | 11 (17.2)  | 19 (24.1)  |
| ISS, mean (SD)      | 30.6 (17.5) | 39.4 (17.2)| 38.1 (15.9)| 38.9 (16.1)| 36.7 (17.2)|
| RTS, mean (SD)      | 5.62 (1.8)  | 2.77 (0.66)| 3.50 (0.80)| 3.34 (0.94)| 3.99 (1.1) |
| GCS                 |             |            |            |            |            |
| mean (SD)           | 10.4 (4.8)  | 3.49 (1.1) | 4.73 (2.8) | 4.59 (2.7) | 5.85 (3.5) |
| median              | 12          | 3          | 3          | 3          | 5          |
| mode                | 15          | 3          | 3          | 3          | 3          |
| Entry SBP, mean (SD)| 75.8 (19.2) | 73.1 (17.7)| 75.6 (17.5)| 73.8 (19.7)| 73.9 (19.1)|
| Entry RR, mean (SD) | 20.1 (7.1)  | 12.0 (7.0) | 16.8 (8.7) | 16.4 (8.8) | 17.3 (8.2) |
| TRISS-Predicted Survival Rate, %, mean (SD) | 66 (35) | 21 (21) | 29 (23) | 27 (23) | 38 (30) |
| Mortality           |             |            |            |            |            |
| Predicted           | 34%         | 79%        | 71%        | 73%        | 62%        |
| Actual              | 35% (73/208)| 71% (29/41)| 59% (35/59)| 63% (40/64)| 53% (42/79)|

Abbreviations: DCLHb, Diaspirin cross-linked hemoglobin; GCS, Glasgow Coma Scale; ISS, Injury Severity Score; MOI, Mechanism of Injury; ns, not significant; RR, Respiratory Rate; RTS, Revised Trauma Score; SBP, Systolic Blood Pressure; TRISS, Trauma and Injury Severity Score.

Figure 2. Distribution of RTS 2-5.99 Patients by TRISS-predicted Survival in the US and EU DCLHb Clinical Trials (n = 74)
Exclusion of GCS <5 patients from the RTS 2-5.99 subgroup caused the removal of 48% of this population, leaving 41 patients eligible for potential inclusion (20% of the total population of 208 patients) (Table 7). Without GCS <5 patients, the RTS (4.73 vs. 3.99) and the GCS (8.41 vs. 5.85) were higher when compared with the complete RTS 2-5.99 subgroup (P < .001). TRISS-predicted survival increased to 46% from 38%, and the actual mortality dropped to 49% from 53% with the removal of the GCS <5 patients (P = ns). Excluding GCS <5 patients caused the removal of 83% of patients from the RTS 1-3.99 subgroup, leaving only seven patients (3.4% of the total population).

In the RTS 2-5.99 subgroup, removal of GCS <5 patients resulted in a 55% reduction in the number of blunt trauma patients and only a 26% reduction in eligible penetrating trauma patients (Table 6). TRISS-predicted and actual mortality rates

| Table 5. Demographics and Clinical Variables by Outcome in RTS 2-5.99 Patients |
|-----------------|-----------------|-----------------|-----------------|-----------------|
|                 | Survivors       | Nonsurvivors    | Total           | P               |
| n (%)           | 37 (46.8)       | 42 (53.2)       | 79 (100)        | —               |
| Age, mean (SD)  | 35.6 (12.4)     | 42.6 (23.1)     | 39.3 (19.2)     | .11             |
| Gender          |                 |                 |                 |                 |
| Male, n (%)     | 29 (78.4)       | 32 (76.2)       | 61 (77.2)       | ns              |
| Female, n (%)   | 8 (21.6)        | 10 (23.8)       | 18 (22.8)       |                 |
| Study Site      |                 |                 |                 |                 |
| US, n (%)       | 22 (59.5)       | 14 (33.3)       | 36 (45.6)       | .020            |
| EU, n (%)       | 15 (40.5)       | 28 (66.7)       | 43 (54.4)       |                 |
| Treatment       |                 |                 |                 |                 |
| DCLHb, n (%)    | 16 (43.2)       | 22 (52.4)       | 38 (48.1)       | ns              |
| NS, n (%)       | 21 (56.8)       | 20 (47.6)       | 41 (51.9)       |                 |
| MOI             |                 |                 |                 |                 |
| Blunt, n (%)    | 21 (56.8)       | 39 (92.9)       | 60 (75.9)       | <.001           |
| Penetrating, n (%) | 16 (43.2) | 3 (7.1)      | 19 (24.1)         |                 |
| ISS, mean (SD)  | 28.1 (15.6)     | 45.1 (14.5)     | 36.7 (17.2)     | <.001           |
| RTS, mean (SD)  | 4.33 (1.1)      | 3.68 (1.0)      | 3.99 (1.1)      | <.008           |
| GCS Score       |                 |                 |                 |                 |
| mean (SD)       | 6.46 (3.9)      | 5.31 (3.1)      | 5.85 (3.5)      | .15             |
| median          | 5               | 3               | 5               |                 |
| mode            | 3               | 3               | 3               |                 |
| Entry SBP, mean (SD) | 72.2 (24.6) | 75.2 (13.9) | 73.9 (19.1) | ns |
| Entry RR, mean (SD) | 17.8 (6.5) | 16.9 (9.3) | 17.3 (8.2) | ns |
| TRISS-Predicted Survival Rate, %, mean (SD) | 54.7 (29.9) | 21.3 (18.2) | 37.6 (29.7) | <.001 |
| Mortality       |                 |                 |                 | 62.4%           |
| Predicted       |                 |                 |                 | 53.2% (42/79)   |
| Actual          |                 |                 |                 | P = ns          |
The population being studied in this RTS as inclusion criterion analysis is significant in that all of the patients who were included in the studies were hypotensive per the entry criteria, and the Glasgow Coma Scale (GCS) score only. The RTS has been studied for its ability to predict outcomes. Specifically, Lichtveld determined that along with advanced age, the triage RTS is the single strongest predictor of in-hospital mortality. However, Giannakopoulos and co-authors found that this same cut-off for hemorrhagic shock studies. This combined patient population is sufficiently large to allow for the study of how RTS ranges of different values could influence the enrollment of patients in future HBOC clinical trials. Additionally, there are no significant differences in demographic variables between these two study populations.

The clinical efficacy trials of DCLHb provide an excellent setting for a retrospective review of the utility of the RTS in attempting to isolate a subgroup of patients with midrange mortality. Combining the two DCLHb clinical trial populations is reasonable because although they come from two different clinical settings, they are otherwise similar to other traumatic hemorrhagic shock studies. This combined patient population was found to be as sensitive as and more specific than the triage-RTS and RTS. Despite extensive study of the RTS as an outcome predictor, it has not been proposed or studied as an entry criterion for any traumatic hemorrhagic shock clinical trials.

The most important clinical trials that have tested hemoglobin-based oxygen carriers for this indication are the two DCLHb clinical trials and the recently published US prehospital clinical trial of PolyHeme, a polymerized hemoglobin solution. The NMRC clinical trial of HBOC-201 RESUS proposed the use of the RTS for the prehospital identification of traumatic hemorrhagic shock patients who might optimally be studied with the ability to identify patients of uniform mortality risk. The RTS was proposed as the inclusion criteria because it can be quickly calculated by paramedics in the prehospital setting with the use of systolic blood pressure (SBP), respiratory rate (RR) and the Glasgow Coma Scale (GCS) score only. The RTS has been studied for its ability to predict outcomes. Specifically, Lichtveld determined that along with advanced age, the triage RTS is the single strongest predictor of in-hospital mortality. However, Giannakopoulos and co-authors found that this same cut-off for trauma patients may not be adequate for helicopter emergency medical services cancellations as it could lead to up to 17% under-triage of major trauma patients. Other scores have been examined for their ability to predict mortality, such as the Mechanism, Glasgow Coma Scale, Age, and Arterial Pressure (MGAP) developed by Sartorius et al. The MGAP was found to be as sensitive as and more specific than the triage-RTS and RTS. Despite extensive study of the RTS as an outcome predictor, it has not been proposed or studied as an entry criterion for any traumatic hemorrhagic shock clinical trials.

Figure 3. Distribution of RTS 2-5.99 Patients without GCS <5 Patients by TRISS-predicted Survival in the US and EU DCLHb Clinical Trials (n = 39)
calculation of the RTS, is the most important variable in the mortality prediction capabilities of different RTS ranges. The only other important attribute of this population is the overall mortality rate, which was comparable to the TRISS-predicted mortality rate. Of note is the fact that the TRISS survival predictions for this 208 patient population also generated a bimodal distribution, with larger numbers of patients with very low and very high mortality risk, similar to Riou’s observations. Because the TRISS-predicted and actual mortality rates were for the most part comparable in these studies, it suggests that the TRISS-prediction model is useful when attempting to identify an optimal patient population for study.

The RTS values, GCS scores, and TRISS predictions were all correlated with mortality, establishing that this dataset could be used in the search for an optimal RTS range that could identify a mid-range mortality risk population for future traumatic hemorrhagic shock studies. The RTS entry criterion range of 1-3.99 generated a subgroup whose mortality rate was twice that of the overall population’s mortality rate. This finding reflects the desire to study a population with a sufficiently high mortality rate to allow for assessment of treatment efficacy. This study examined RTS ranges that were higher (2-4.99), wider (1-4.99), and higher and wider in value (2-5.99). There was no scientific basis for the selection of these specific RTS strata other than the assumption that higher and wider RTS strata ranges would identify more patients with a uniform, mid-range mortality risk, as well as a larger number of patients eligible for study.

The RTS range of 2-5.99 was studied further because of its optimal mortality risk, which included a 62% predicted and 53% actual mortality rates. This population also included the largest

### Table 6. Demographics and Clinical Variables by Mechanism of Injury in RTS 2-5.99 Patients

| Abbreviations: GCS, Glasgow Coma Scale; ISS, Injury Severity Score; MOI, Mechanism of Injury; ns, not significant; RR, Respiratory Rate; RTS, Revised Trauma Score; SBP, Systolic Blood Pressure; TRISS, Trauma and Injury Severity Score. |

| Including GCS <5 | Excluding GCS <5 |
|------------------|------------------|
| **Blunt** | **Penetrating** | **P** | **Blunt** | **Penetrating** | **P** |
| n (%) | 60 (75.9) | 19 (24.1) | — | 27 (65.9) | 14 (34.1) | — |
| Age, mean (SD) | 40.3 (20.9) | 36.2 (12.4) | ns | 42.0 (20.0) | 36.0 (13.8) | ns |
| Gender | | | | | | |
| Male, n (%) | 15 (78.9) | 46 (76.7) | ns | 15 (78.9) | 46 (76.7) | ns |
| Female, n (%) | 4 (21.1) | 14 (23.3) | ns | 4 (21.1) | 14 (23.3) | ns |
| Study Site | | | | | | |
| US, n (%) | 26 (43.3) | 10 (52.6) | ns | 8 (29.6) | 5 (35.7) | ns |
| EU, n (%) | 34 (56.7) | 9 (47.4) | ns | 19 (70.4) | 9 (64.3) | ns |
| ISS, mean (SD) | 42.1 (14.7) | 20.9 (14.3) | <.001 | 45.5 (15.4) | 23.1 (14.9) | <.001 |
| RTS, mean (SD) | 3.77 (1.1) | 4.67 (0.94) | <.001 | 4.57 (0.87) | 5.02 (0.70) | ns |
| GCS Score | | | | | | |
| mean (SD) | 5.20 (3.1) | 7.89 (4.0) | .003 | 7.85 (2.9) | 9.50 (3.3) | ns |
| median | 3 | 7 | 7 | 7 |
| mode | 3 | 7 | 5 | 7 |
| Entry SBP, mean (SD) | 74.6 (16.4) | 71.1 (28.6) | ns | 67.5 (17.9) | 60.7 (25.8) | ns |
| Entry RR, mean (SD) | 16.9 (5.8) | 17.4 (8.9) | ns | 20.9 (8.6) | 16.9 (5.8) | ns |
| TRISS-Predicted Survival Rate, %, mean (SD) | 29.0 (25.8) | 62.4 (26.5) | <.001 | 33.7 (28.8) | 66.0 (26.8) | <.002 |
| Mortality | | | | | | |
| Predicted | 71.0% | 37.6% | 66.3% | 34.0% |
| Actual | 65.0% (39/60) | 15.8% (3/19)* | 66.7% (18/27) | 14.3% (2/14)* |
| P = ns | P = ns | P = ns | P = ns |

*P = ns for Blunt vs. Penetrating.
percentage of patients and the largest proportion of penetrating trauma patients as compared to other strata that were examined. Of note in the RTS 2-5.99 population is the observation that the mean GCS was 5.8 and the median GCS was five. Although the GCS scores for these RTS 2-5.99 patients suggest significant traumatic brain injury, this is still the preferred stratum because all of the other proposed RTS strata in this study had a median GCS score of three, which signifies the potential inclusion of a greater numbers of GCS <5 patients, who are known to undermine the efficacy detection capabilities of traumatic hemorrhagic shock clinical trials. The proposed optimal RTS 2-5.99 range was compared to the initially proposed RTS 1-3.99 range with the finding that the TRISS-predicted survival in the RTS 2-5.99 stratum was nearly two times higher as well as much closer to the desired 50% mortality risk.

The RTS 2-5.99 patient subgroup was analyzed based on mechanism of injury in order to determine if this stratum might be more useful with blunt or penetrating trauma patients. A difference would suggest the need for separate traumatic hemorrhagic shock clinical trials based on mechanism of injury. Blunt trauma patients in the RTS 2-5.99 stratum were found to be more severely injured than the penetrating trauma patients. Consequentially, blunt trauma RTS 2-5.99 patients had a four-fold higher mortality rate as compared to the penetrating trauma patients in the same RTS range. Despite the outcome difference in patients with blunt and penetrating trauma mechanisms, the overall conclusions from this study regarding the use of the RTS are still sound. However, any future studies of how the RTS might be used as an entry criterion for traumatic hemorrhagic shock clinical trials should optimally analyze blunt and penetrating injury patients independently.

### Table 7. Patient Demographics and Clinical Variables Including and Excluding GCS <5

|                | Including GCS <5 | Excluding GCS <5 | P       | Including GCS <5 | Excluding GCS <5 | P       |
|----------------|------------------|------------------|---------|------------------|------------------|---------|
| n (%)          | 79 (65.8)        | 41 (34.2)        | —       | 41 (85.4)        | 7 (14.6)         | —       |
| Age, mean (SD) | 39.3 (19.2)      | 39.9 (18.1)      | ns      | 38.4 (20.0)      | 31.6 (16.6)      | ns      |
| Gender         |                  |                  |         |                  |                  |         |
| Male, n (%)    | 61 (77.2)        | 33 (80.5)        | ns      | 30 (73.2)        | 5 (71.4)         | ns      |
| Female, n (%)  | 18 (22.8)        | 8 (19.5)         |         | 11 (26.8)        | 2 (28.6)         |         |
| MOI            |                  |                  |         |                  |                  |         |
| Blunt, n (%)   | 60 (75.9)        | 27 (65.9)        | .15     | 37 (90.2)        | 6 (85.7)         | ns      |
| Penetrating, n (%) | 19 (24.1)    | 14 (34.1)        |         | 4 (9.8)          | 1 (14.3)         |         |
| ISS, mean (SD) | 36.7 (17.2)      | 37.0 (18.6)      | ns      | 39.4 (17.2)      | 40.1 (21.6)      | ns      |
| RTS, mean (SD) | 3.99 (1.1)       | 4.73 (0.84)      | <.001   | 2.77 (0.66)      | 3.41 (0.53)      | .024    |
| GCS Score      |                  |                  |         |                  |                  |         |
| mean (SD)      | 5.85 (3.7)       | 8.41 (3.12)      | <.001   | 3.49 (1.1)       | 5.50 (1.31)      | <.001   |
| median         | 5                | 7.5              |         | 3                | 5                |         |
| mode           | 3                | 5                |         | 3                | 5                |         |
| Entry SBP, mean (SD) | 73.9 (19.1) | 65.8 (20.0)      | .055    | 73.1 (17.7)      | 72.9 (15.6)      | ns      |
| Entry RR, mean (SD) | 17.3 (8.2)   | 19.4 (7.9)       | ns      | 12.0 (7.0)       | 14.2 (7.7)       | ns      |
| TRISS-Predicted Survival Rate, %, mean (SD) | 37.6 (29.7) | 46.2 (31.9) | ns | 20.9 (20.7) | 34.1 (30.3) | ns |
| Mortality      |                  |                  |         |                  |                  |         |
| Predicted      | 62.4%            | 53.8%            |         | 79.1%           | 65.9%            |         |
| Actual         | 53.2% (42/79)    | 48.8% (20/41)   | a       | 70.7% (29/41)   | 57.1% (4/7)      | a       |

\[ P = \text{ns} \] for Including GCS <5 vs. Excluding GCS <5.

Abbreviations: GCS, Glasgow Coma Scale; ISS, Injury Severity Score; MOI, Mechanism of Injury; ns, not significant; RR, Respiratory Rate; RTS, Revised Trauma Score; SBP, Systolic Blood Pressure; TRISS, Trauma and Injury Severity Score.

\[ ^a \text{P} = \text{ns} \] for Including GCS <5 vs. Excluding GCS <5.
When analyzing the RTS 2-5.99 patients with regard to treatment group, there were no observed differences with regard to RTS, GCS, ISS, SBP, RR, TRISS-predicted survival, or actual mortality rates. As such, DCLHb treatment effects did not render the RTS entry criteria stratum analysis invalid.

Because it is known that GCS 5 patients might best be excluded from traumatic hemorrhagic shock clinical trials due to the potential for including severely head-injured patients, and because the recently completed PolyHeme study excluded GCS 3 and 4 patients, and because the RESUS study intended to exclude these patients, further analysis examined both RTS strata with the exclusion of GCS 5 patients. Excluding GCS 5 victims resulted in a remaining population that included half of the RTS 2-5.99 population (one-fifth of the overall 208 patient population). This exclusion also was noteworthy because although it removed over half of the blunt trauma patients, it only removed around a quarter of the penetrating trauma patients in this stratum, leading to a more equal distribution of blunt and penetrating trauma patients in the RTS 2-5.99 stratum (66% and 34%, respectively). The GCS characteristics of this refined RTS 2-5.99 population also were enhanced in that the mean GCS increased to 8.4, and the median GCS was in the 7-8 range. As such, the RTS 2-5.99 stratum without GCS 5 patients perhaps optimally identifies trauma patients whose GCS score suggests altered mental status as a result of traumatic hemorrhagic shock as opposed to the presence of concomitant severe traumatic brain injury. When the GCS 5 patients are excluded, the mortality rate of the RTS 2-5.99 stratum approaches the desired midrange mortality risk, with predicted TRISS survival of 46% and actual mortality of 49%.

When the RTS 2-5.99 stratum without the GCS 5 patients was analyzed based on mechanism of injury, it was noted that the majority of excluded patients came from the blunt subgroup. This reflects the fact that most of the GCS <5 patients were blunt trauma patients with concomitant significant closed head injury. Once these GCS <5 patients were removed, the RTS and GCS were comparable in both blunt and penetrating subgroups; this finding suggests that a study utilizing RTS 2-5.99 patients would not necessarily require two separate studies, given the similar RTS and GCS characteristics of the blunt and penetrating subgroups.

Nearly half of the patients in the RTS 2-5.99 stratum without GCS 5 had a predicted survival between 20 and 80% which is similar to the range observed in the RTS 2-5.99 subgroup that included GCS 5 patients. This stable TRISS-predicted survival rate suggests that the RTS 2-5.99 stratum selects a patient population with mid-range mortality risk, even if patients with severe traumatic brain injury are excluded. The overall distribution of patient in the RTS 2-5.99 stratum without GCS 5 patients demonstrates nearly one quarter of patients in each of the extremely low and extremely high survival risk groups, a finding that suggests that this distribution may be the best that can be achieved when conducting a traumatic hemorrhagic shock study with an undifferentiated population of trauma patients.

The ROC data showed that GCS was similar to RTS in its ability to predict mortality in the traumatic hemorrhagic shock patient, and that the SBP and RR values were not adequately predictive of outcome. This finding does not suggest that SBP is not a useful entry criterion for a traumatic hemorrhagic shock study; rather SBP simply is not predictive of outcome because most of the patients from these two trauma studies had similar hypotensive SBPs at the time of study entry per the inclusion criteria. When using the RTS as an entry criterion, patients will be enrolled whose shock compensatory mechanisms create a similar physiologic state, as measured by GCS, SBP, and RR.
and whose TRISS-predicted survival is related primarily to injury severity. When the GCS <5 patients were removed and the 28-day mortality ROC analysis was repeated, it was observed that the sensitivity and specificity of each component and the composite RTS decreased slightly, with GCS and RTSs remaining the best predictors of mortality. This drop in predictive power with the removal of GCS < 5 patients indicates the importance of GCS scores of less than five in predicting 28-day mortality.

Limitations

One limitation of this study is the small patient population (208) that resulted from the early termination of the DCLHb efficacy trials. If a larger trauma population from an HBOC clinical trial such as the PolyHeme clinical trial was used for this type of analysis, it might be possible to generalize this data to a larger population of trauma patients who could be enrolled in future traumatic hemorrhagic shock clinical trials. Another possible limitation is the fact that the RTS strata were chosen empirically and not based on previously established optimal RTS strata. This reflects the fact that there is limited data relating RTS to mortality prediction at the time of initial medical contact, aside from the Lichtveld data, which does not specifically identify clinically useful strata.24

Future research should examine what could be optimal entry criteria using larger traumatic hemorrhagic shock clinical trial databases in order to establish what might be the optimal RTS range, and whether the use of the RTS as an entry criterion is actually better than the traditional use of SBP < 90 mm Hg as a measure of critical hemorrhage and the need for optimal traumatic hemorrhagic shock resuscitation. Additional work should examine the effect of excluding patients with low GCS scores on the conduct of traumatic hemorrhagic shock clinical trials, both with regard to what type of patients will be studied and how the exclusion of these patients will affect patient enrollment and the ultimate feasibility of the clinical trial.

Conclusion

Based on study analysis of the DCLHb clinical trial data, an RTS range of 2-5.99 is a more appropriate range for the identification of traumatic hemorrhagic shock patients whose injury severity is more closely associated with a mortality risk near 50% than that observed in the RTS 1-3.99 stratum. The RTS 2-5.99 subgroup allows for identification of patients whose survival distribution may be optimal for study because half of the patients have an intermediate mortality risk. The exclusion of GCS < 5 patients refines the RTS 2-5.99 patient population by creating a smaller subgroup with higher GCS scores, moderate injury severity, and TRISS-predicted and actual mortality rates that are nearly identical to the desired 50% mortality risk that could optimally demonstrate the effects of a novel traumatic hemorrhagic shock therapy such as a hemoglobin solution. When studying RTS 2-5.99 patients with exclusion of GCS < 5 patients, the similar distribution and predictive characteristics of blunt and penetrating patients may allow for both to be studied in one traumatic hemorrhagic shock clinical trial. Because the GCS score showed mortality prediction comparable to the RTS, any future clinical trial must consider the GCS score as an independent inclusion or exclusion criterion.

References

1. Krausz MM. Controversies in shock research: hypertonic resuscitation—pros and cons. Shock. 1995;3(3):69-72.
2. Sauer A, Moore FA, Moore EE, et al. Epidemiology of trauma deaths: a reassessment. J Trauma. 1995;48(2):185-193.
3. Acosta JA, Yang JC, Winchell RJ, et al. Lethal injuries and time to death in a level I trauma center. J Am Coll Surg. 1998;186(5):528-533.
4. Kaurav DS, Wade CE. The epidemiology and modern management of traumatic hemorrhage: US and international perspectives. Crit Care. 2005;9 Suppl 5:S1-9.
5. Duranteau J, Harois A. Hemorrhagic shock [in French]. Rev Prat. 2006;56(8):849-857.
6. Rossaint R, Ceryn U, Coats TJ, et al. Key issues in advanced bleeding care in trauma. Shock. 2006;26(4):322-331.
7. Jaeh JS, Walker V, Manoochehrl K. Blood substitutes as pharmacotherapies in clinical practices. Curr Opin Anaesthesiol. 2007;20(4):325-330.
8. Winslow RM. Cell-free oxygen carriers: scientific foundations, clinical development, and new directions. Biochim Biophys Acta. 2008;1784(10):1382-1386.
9. Moore EE, Johnson JL, Moore FA, Moore HB. The USA Multicenter Prehospital Hemoglobin-based Oxygen Carrier Resuscitation Trial: scientific rationale, study design, and results. Crit Care Clin. 2009;25(2):325-356. Table of Contents.
10. Freilich D, Pearce LB, Pitman A, et al. HBOC-201 vasoactivity in a phase III clinical trial in orthopedic surgery subjects: extrapolation of potential risk for acute trauma trials. J Trauma. 2009;66(2):365-376.
11. Roui B, Landais P, Vivien B, et al. Distribution of the probability of survival is a strategic issue for randomized trials in critically ill patients. Anesth Analg. 2001;93(1):56-63.
12. Dellingner RP, Vincent JL, Marshall J, Reinhart K. Important issues in the design and reporting of clinical trials in severe sepsis and acute lung injury. J Crit Care. 2008;23(4):493-499.
13. George SL. Statistical issues in translational cancer research. Clin Cancer Res. 2008;14(19):5954-5958.
14. Kerner T, Ahlers O, Veit S, Riou B, Saunders M, Pison U. DCL-Hb for trauma patients with severe hemorrhagic shock: the European “On-Scene” multicenter study. Intensive Care Med. 2003;29(3):378-385.
15. Moore L, Lavoie A, Turgeon AF, et al. The trauma risk adjustment model: a new model for predicting trauma mortality. J Trauma. 1998;23(4):759-761.
16. Mongan PD, Moon-Massat PF, Rentko V, Mihok S, Dragovich A, Sharma P. Regional blood flow after serial normovolemic exchange transfusion with HBOC-201 (Hemopure) in anesthetized swine. J Trauma. 2009;67(1):51-60.
17. Freilich D, NMRC. Review of a Proposed Clinical Trial of HBOC-201 in Trauma. December 14, 2006.
18. Marinaro J, Smith J, Tawil I, Billstrand M, Crookston KP. HBOC-201 use in traumatic brain injury: case report and review of literature. Transfusion. 2009;49(10):2054-2059.
19. Jhrl JS, Mackenzie C, Pearce LB, Pitman A, Greenburg AG. HBOC-201 as an alternative to blood transfusion: efficacy and safety evaluation in a multicenter phase III trial in elective orthopedic surgery. J Trauma. 2008;66(4):1484-1497.
20. Champion HR, Sacco WJ, Copes WS, Gann DS, Genereux TA, Planegage M. A revision of the Trauma Score. J Trauma. 1989;29(3):623-629.
21. Gabbe BJ, Cameron PA, Finch CF. Is the revised trauma score still useful? ANZ J Surg. 2003;73(11):944-948.
22. Guzzo JL, Bochicchio GV, Napolitano LM, Malone DL, Meyer W, SCALE TM. Prediction of outcomes in trauma: anatomic or physiologic parameters? J Am Coll Surg. 2005;201(6):891-897.
23. Esme H, Solak O, Yurumes Y, et al. The prognostic importance of trauma scoring systems for blunt thoracic trauma. Thorax Cardiovasc Surg. 2007;55(3):190-195.
24. Lichtveld RA, Spijkers AT, Hoogendoorn JM, Panhuizen IF, van der Werken C. Triage Revised Trauma Score change between first assessment and arrival at the hospital to predict mortality. Int J Emerg Med. 2008;1(1):21-26.
25. Ahmad HN. Evaluation of revised trauma score in polytraumatized patients. J Cell Physiol Surg Pathol. 2004;14(5):286-289.
26. Sloan EP, Koenigsberg M, Gen D, et al. Diaspiron cross-linked hemoglobin (DCLHb) in the treatment of severe traumatic hemorrhagic shock: a randomized controlled efficacy trial. JAMA. 1999;282(19):1857-1864.
27. Boyd CR, Tolson MA, Copes WS. Evaluating trauma care: the TRISS method. Trauma Score and the Injury Severity Score. J Trauma. 1987;27(4):370-378.
28. Moore EE, Johnson JL, Cheng AM, Masuno T, Bunerje A. Insights from studies of blood substitutes in trauma. Shock. 2005;24(4):197-205.
29. Carmichael FJ, Ali AC, Campbell JA, et al. A phase I study of oxidized raffinose from Rhizopus oryzae as an alternative to blood transfusion. Transfusion. 2006;46(4):623-630.
30. Alayash AI, D’Agnillo F, Buehler PW. First-generation blood substitutes: what have we learned? Biochemical and physiological perspectives. Expert Opin Biol Ther. 2007;7(5):653-675.
31. Pedron L, Peck N, de Jonge E, Scheffer GJ, de Keizer NF. The use of a registry database in clinical trial design: assessing the influence of entry criteria on statistical power and number of eligible patients. Int J Med Inform. 2007;76(2-3):176-183.
32. Atrosi I, McCabe SJ. The early steps in clinical research: importance of entry criteria and true randomization. J Hand Surg Am. 1998;23(4):759-761.
Appendix 1: United States (US) DCLHb Clinical Efficacy Trial

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Appendix 2: European Union (EU) DCLHb HOST Clinical Efficacy Trial

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- Docteur Guerrini, MD (Principal Investigator, Hopital A. Mignot, Le Chesnay)
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**August 2012**

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