Review of meta-analyses evaluating surrogate endpoints for overall survival in oncology

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Abstract: Overall survival (OS) is the gold standard in measuring the treatment effect of new drug therapies for cancer. However, practical factors may preclude the collection of unconfounded OS data, and surrogate endpoints are often used instead. Meta-analyses have been widely used for the validation of surrogate endpoints, specifically in oncology. This research reviewed published meta-analyses on the types of surrogate measures used in oncology studies and examined the extent of correlation between surrogate endpoints and OS for different cancer types. A search was conducted in October 2010 to compile available published evidence in the English language for the validation of disease progression-related endpoints as surrogates of OS, based on meta-analyses. We summarize published meta-analyses that quantified the correlation between progression-based endpoints and OS for multiple advanced solid-tumor types. We also discuss issues that affect the interpretation of these findings. Progression-free survival is the most commonly used surrogate measure in studies of advanced solid tumors, and correlation with OS is reported for a limited number of cancer types. Given the increased use of crossover in trials and the availability of second-/third-line treatment options available to patients after progression, it will become increasingly more difficult to establish correlation between effects on progression-free survival and OS in additional tumor types.

Keywords: progression endpoints, correlation, cancer

Introduction
Rapid changes in our understanding of cancer biology and genetics, accompanied by the advent of newer targeted agents, are affecting every level of drug development, including molecule screening, development planning, study designs, regulatory decision making, and reimbursement choices. Although overall survival (OS) remains the gold standard for assessing patient benefit from new drug therapies for cancer, practical factors may preclude the collection of unconfounded OS data. Showing a survival advantage of one treatment over another in cancer clinical trials can take years, and if patients take other treatments that improve survival after disease progression, attributing benefits confidently to a single agent or designing a feasible trial protocol with enough patients and duration of follow-up may not be possible. In addition, the length of survival post-progression may make it difficult to detect a survival advantage, even if one exists, due to the random variation associated with patient heterogeneity and the influences of subsequent therapy.1 An obvious need exists for well-defined and valid measures of benefit from anticancer treatment that can be assessed earlier in the course of the disease than patient death. Since approval and access to a new product hinges on successful Phase 3 clinical trial results, surrogate endpoints that
could support earlier decision making would provide patients with new treatments sooner and reduce the costs of drug development, as has been seen in many other therapy areas (eg, HIV/AIDS and cardiovascular disease).

A surrogate endpoint in a clinical trial is “a substitute for a clinically meaningful endpoint that measures directly how a patient feels, functions or survives.” A surrogate endpoint must be clinically relevant, sensitive to treatment, and measurable. Surrogates are particularly valuable for drug development in diseases where increased patient survival is the goal of treatment, but a long time is required to observe this endpoint directly. For example, in studies of antihypertensives, blood pressure reduction is generally accepted as a surrogate endpoint for the reduction of longer-term and more severe cardiovascular endpoints. In general, justification for the use of a surrogate depends on multiple considerations that vary depending on the disease or specific cancer, drug mechanism of action, phase of development, patient subgroup, and availability of alternate treatments. For example, response rate has a role in evaluating the antitumor activity of new drugs in Phase 1 and 2 studies, but it is not recognized as an endpoint showing patient benefit in all tumors. This distinction is partly based on the fact that the benefit of a partial tumor response is not necessarily outweighed by the toxicity associated with treatment; also, the proportion of patients responding is not always a valid predictor of survival or other clinical benefits. Time to progression (TTP), an endpoint that evaluates disease progression but censors deaths rather than counting them as events, has fallen out of favor in contemporary Phase 3 trials. Progression-free survival (PFS) is considered a more realistic assessment of treatment efficacy, since it counts both progression and deaths as part of the endpoint.

For any stage in the drug development process, use of a surrogate endpoint rather than the target endpoint may shorten clinical trials but increase the chance of false positive results. Validation of surrogate endpoints is typically based on the Prentice criterion, a set of conditions that specify the relationship between the treatment and endpoints under consideration. Changes in a surrogate endpoint that are induced by a therapy are expected to reflect changes in a clinically meaningful target endpoint.

During the past decade, a body of work has developed that uses meta-analytic techniques to investigate progression-related endpoints as possible surrogates for overall survival in patients with solid tumors. The meta-analyses conducted for surrogate endpoint validation in oncology are somewhat atypical in that the objective is to establish the relationships between endpoints, rather than summarizing treatment effects on a single endpoint. To accomplish this goal, investigators use a technique called meta-regression to model a treatment effect for survival against a treatment effect for the potential surrogate endpoint. For example, based on individual-patient data or summary data from multiple clinical trials, the hazard ratio (HR) for comparing two treatments on overall survival (HRos) can be regressed on the hazard ratio for PFS (HRpfs), resulting in an equation such as the following:

\[ HR_{os} = \mu + (\beta \times HR_{pfs}) + \epsilon \]  

where \( \mu \) represents an intercept, \( \beta \) is the slope of the line showing the linear relationship of the hazard ratios, and \( \epsilon \) is the unexplained variance. In Equation (1), each study contributes one observation, typically weighted by the variance of the study-specific HR. Such an analysis expresses the relationship between differences in effect sizes for progression and survival across multiple trials and gives an idea of how strongly the endpoints are linked mathematically, assuming a linear relationship.

In other words, the meta-regression equation shows the predicted relationship between the hazard ratios for progression-free survival and overall survival, based on the studies included. If the slope (\( \beta \)) of this equation equals 1, assuming a negligible intercept, the treatment effects on survival are expected to be of similar magnitude to effects on PFS. Models may address covariates or factors that can influence the endpoint relationship, and sometimes the meta-analysis is repeated on different patient subgroups or subsets of studies. Meta-regression equations take many different forms in the published literature, depending on factors such as which endpoint was evaluated, whether a transformation (logarithm) was used, what statistical model was implemented, and how study weights were derived. Some authors model the difference between treatments in median months to the event as the treatment effect, or analyze data from the study arms separately.

Typically, authors present the simple correlation \( r \) between the treatment effect measures across trials. Correlation values are close to one if the treatment effects tend to go in the same direction. In other words, correlation is high if the hazard ratios for PFS and OS are similar across trials; correlation is low if the hazard ratios are unrelated or in opposite directions. A related measure (\( R^2 \) or R-squared) is derived from the meta-regression equation to indicate how much variance
in OS is explained by the potential surrogate PFS. In the very simplest case, \( R^2 \) is equal to a squared correlation estimate (ie, \( R^2 = r \times r \)). Some authors denote whether \( R^2 \) is based on (a) models of median OS and median PFS from individual study arms (\( R^2_{\text{ind}} \)) or (b) models of hazard ratio for OS and hazard ratio for PFS (\( R^2_{\text{trial}} \)) (Figure 1). Otherwise, the publication may simply show \( r \) or \( R^2 \) and leave it to the reader to distinguish which one is being used. Because \( R^2_{\text{ind}} \) addresses whether treatment effects on the surrogate endpoint are associated with treatment effects on survival, it is particularly important for drug development.

While many authors have presented meta-analyses of the relationship between treatment effects on PFS and OS in specific cancers, we have not seen a compilation of this evidence across tumor types. In October 2010, we reviewed published meta-analyses used for the validation of progression-related surrogate endpoints in multiple tumor types with the goal of summarizing the existing evidence, assessing differences and similarities across tumor types, and elucidating the assorted challenges associated with this topic. We conducted a comprehensive and iterative review of the literature, identifying more than 1000 titles and scanning more than 100 full-text articles for relevance, including review articles, methods papers, and commentary from clinicians, statisticians, regulators, and payers. We focused on evidence on disease-progression endpoints that was likely to be used in Phase 3 studies conducted for regulatory approval of new treatments for patients with solid tumors, especially in the advanced/metastatic setting. Relevant papers that were published while this manuscript was in preparation were also incorporated.

Multiple solid tumors were represented in publications (Table 1) that were identified as using meta-analyses to evaluate the following progression-related endpoints as surrogates for OS:
- Disease-free survival – adjuvant setting only – or event-free survival
- Response rate or objective response rate
- TTP
- PFS

Many articles presented a meta-regression showing the relationship between treatment effects on these endpoints and treatment effects on OS, but we found little consistency in the model form or presentation of predicted values.

**Summary of available evidence by cancer type**

**Colorectal cancer**

Colorectal cancer is currently the only metastatic tumor type for which consistently strong validation evidence is available on progression-free survival as a surrogate for overall survival. A number of published papers that examined progression-related endpoints as potential surrogates for survival in colorectal cancer presented graphs with regression equations for the relationship between treatment effects on the surrogate endpoint and on survival. Studies with results falling almost perfectly on a diagonal line to depict the relationship between hazard ratios for disease-free survival and overall survival contributed to the acceptance of this endpoint as a valid surrogate endpoint in the adjuvant treatment setting for colorectal cancer.\(^{13,14}\)

The published meta-analyses provide strong evidence for the correspondence between progression-free survival and overall survival in the metastatic disease setting, although the equations predict that hazard ratios for overall survival will be somewhat closer to the null than hazard ratios for progression-free survival. Buyse and colleagues\(^{15}\) concluded that a novel therapy producing a 10% risk reduction for progression-free survival would yield an estimated 5.4% improvement in overall survival. Buyse and colleagues\(^{16}\) predicted that risk reductions in colorectal cancer would be lower on overall survival than on progression-free survival, and they suggested a threshold effect that a new treatment would have to show for the hazard ratio for

![Figure 1](https://www.dovepress.com/figures/diagram-of-the-relationship-between-treatment-and-pfs-as-potential-surrogate-endpoint-for-os.jpg)

**Figure 1** Diagram of the relationship between treatment and PFS as potential surrogate endpoint for OS. **Abbreviations:** OS, overall survival; \( R^2_{\text{ind}} \), R-squared from meta-regression of median PFS and OS from individual treatment arms; \( R^2_{\text{trial}} \), R-squared from meta-regression of hazard ratios from each trial.
Table 1  Meta-analytic evidence for correlation between progression endpoints and OS in multiple solid tumors as reported in publications

| Cancer setting | Surrogate | Correlation between surrogate and OS endpoints | Correlation between treatment effects on surrogate and on OS | References |
|----------------|-----------|-----------------------------------------------|----------------------------------------------------------|------------|
| A              |           |                                               |                                                          |            |
| Colon, adjuvant| DFS       | R: 0.88; R²: 0.85                            | R: 0.94; R²: 0.90                                         | 9, 13      |
| Colorectal, adjuvant | DFS | R: 0.95                                           | R: 0.94; R²: 0.90                                         | 14         |
| Colorectal, advanced | Response rate | R: 0.408                                    | R: 0.38                                                   | 56         |
| Colorectal, metastatic | Response rate | R: 0.59                                      | R: 0.10                                                   | 15         |
| Colorectal, metastatic | Response rate | R: 0.59                                      | R: 0.33                                                   | 29         |
| Colorectal, metastatic | TTP       | R: 0.24                                       | R: 0.33                                                   | 15         |
| Colorectal, advanced | PFS       | R: 0.82                                       | R: 0.99; R: 0.74 (sensitivity analysis)                   | 16         |
| Colorectal, metastatic | PFS       | R: 0.79                                       | Difference in PFS R²: 0.65                                | 15         |
| Colorectal, metastatic | PFS       | R: 0.481                                      | R²: 0.10                                                  | 17         |
| Colon or colorectal, metastatic | PFS | R²: 0.38 (only DFS), 0.39 (node ±), 0.37 (hormone trials), 0.43 (chemotherapy trials) | R: 0.38; R² trial: 0.64; HR ORR R²: 0.20; Difference in ORR R²: 0.20; HR ORR R²: 0.10 | 22         |
| B              |           |                                               |                                                          |            |
| Ovarian, advanced | TTP       | R²: 0.88                                      | R: 0.94                                                   | 58         |
| Ovarian, metastatic | PFS       | R²: 0.73                                      | Difference in PFS R²: 0.60; HR PFS R²: 0.73               | 57         |
| Ovarian: platinum-resistant | PFS | At 6 months, R: 0.66                        | R²: 0.95                                                  | 59         |
| Ovarian: advanced | PFS       | R²: 0.70                                      | R: 0.95                                                   | 60         |
| C              |           |                                               |                                                          |            |
| Breast, adjuvant | DFS       | R²: 0.38 (only DFS), 0.39 (node ±), 0.37 (hormone trials), 0.43 (chemotherapy trials) | R²: 0.38; R² trial: 0.64; HR ORR R²: 0.20; Difference in ORR R²: 0.20; HR ORR R²: 0.10 | 22         |
| Breast, metastatic | ORR       |                                               |                                                          | 61         |
| Breast, metastatic | Response rate |                                               | R: 0.57                                                   | 62         |
| Breast, metastatic | TTP       | R: 0.682                                      | R: 0.49                                                   | 62         |
| Breast, advanced | TTP       | R²: 0.67 for trials before 1990               | R: 0.41 for trials after 1990                             | 25         |
| Breast, advanced | PFS       | Anthracyclines, R²: 0.49                      | Taxanes, R²: 0.35                                         | 24         |
| Breast, metastatic | PFS       | R: 0.688                                      | R: 0.48                                                   | 62         |
| Breast, metastatic | TTP, PFS  | TTP and PFS: R²: 0.38; Overall R²: 0.30; HR PFS R²: 0.52; Anthracyclines R²: 0.43; hormonal R²: 0.24; HER2+ R²: 0.93 | R²: 0.48; R² trial: 0.64; HR PFS R²: 0.20; Difference in ORR R²: 0.20; HR ORR R²: 0.10 | 23         |
| D              |           |                                               |                                                          |            |
| NSCLC, advanced | Response rate |                                               | R²: 0.16                                                  | 29         |
| NSCLC, advanced | Response rate |                                               | R²: 0.16                                                  | 63         |
| NSCLC          | TTP       | R²: 0.19                                      |                                                           | 29         |
| SCLC           | PFS       | R²: 0.79                                      |                                                           | 28         |
| E              |           |                                               |                                                          |            |
| Brain (glioblastoma multiforme) | PFS | Kappa statistics: 0.48 to 0.52 | R: 0.53                                                   | 39         |
| Head and neck, locally advanced (radiotherapy trials) | EFS | R: 0.86                                      |                                                           | 38         |
| Prostate, castrate-resistant | PFS | Association 0.30                           |                                                           | 37         |
| Prostate, advanced | PFS       |                                               |                                                           | 8          |
| Renal cell carcinoma | PFS       |                                               |                                                           | 34         |
| Multiple metastatic solid tumors: breast, pancreatic, colon or colorectal, ovarian, renal cell carcinoma, esophago-gastric | PFS | R²: 0.49 HR PFS R²: 0.62 |                                                          | 57         |

Notes: *Author noted significant relationship (P < 0.05) but did not provide R or R²; †looked at proportion with progression at 6 months associated with overall survival at 12 months.

Abbreviations: OS, overall survival; DFS®, disease-free survival; EFS, event-free survival; NSCLC, non-small cell lung cancer; ORR, objective response rate; PFS, progression-free survival; SCLC, small cell lung cancer; TTP, time to progression; R², R-squared.
progression-free survival (HR_{PFS}) to have a beneficial impact on overall survival.

The evidence led to a consensus among researchers that these endpoints were valid surrogates for survival in studies of colorectal cancer.\textsuperscript{15-17} On the other hand, analyses referenced in Table 1A also helped to establish that response rate and time to progression were insufficient as surrogate endpoints for survival in colorectal cancer studies.

**Ovarian cancer**

As long ago as 1992, Torri et al\textsuperscript{18} presented a correlation of endpoints in advanced ovarian cancer by separate treatment arms. The evidence we compiled recommends progression-free survival or even the time to progression as a surrogate (Table 1B). However, researchers have found the data supporting the validity of progression-free survival as a surrogate for overall survival in second- and third-line therapy to be less clear than those for first-line therapy.\textsuperscript{19,20}

Molenberghs and colleagues\textsuperscript{21} compared various surrogacy criteria and concluded that progression-free survival was not a useful endpoint in ovarian cancer because it could take a long time for ovarian cancer to cause symptoms or to be detected by physical examination or imaging studies, and that progression was typically followed by death within a few months. Bast and colleagues\textsuperscript{19} summarized expert commentary from a Food and Drug Administration (FDA) workshop on evaluating potential endpoints in ovarian cancer drug research:

PFS seems to correlate with OS, especially when a large effect on PFS is seen. A small increase in PFS may not correlate with OS, however, and crossover confounds the measurement of OS.

However, PFS has gained wide acceptance as a suitable surrogate endpoint in Phase 3 studies for the first-line treatment of metastatic ovarian cancer. Further evidence is needed to support its use in resistant or refractory disease.

**Breast cancer**

Findings from meta-regressions of breast cancer studies are less convincing than those of other tumors; two authors provided predictions indicating that the treatment effect on overall survival was expected to be much smaller than on the progression endpoint in trials of breast cancer, whether in the adjuvant\textsuperscript{22} or advanced setting.\textsuperscript{23} Correlations (R^2_{ind}) between progression-free survival or time-to-progression endpoints and survival ranged from 0.38 to 0.68 when individual treatment arms were analyzed (Table 1C). The association of treatment effects (R^2_{total}) on these endpoints varies widely, ranging from 0.24 to 0.78, and it is not clear whether the variability is due to different treatment types, line of therapy, frequency of crossover or use of other therapies after progression, data quality, analytical approach, or other factors.

Several of the published meta-analyses of surrogacy in breast cancer present analyses of subgroups, in which authors investigated whether a stronger relationship between endpoints may be apparent in certain trial settings. Results seem to vary by which drug class was studied and whether individual-patient data are available for analysis; in other words, the subgroup analyses have generally uncovered mixed information, making it difficult to draw definite conclusions. Miksad and colleagues\textsuperscript{24} reported regression equations that implied a stronger relationship between endpoints for studies using taxane-based rather than anthracycline-based chemotherapies, so in studies of breast cancer, the specific treatment evaluated may affect the correlation between a surrogate endpoint and overall survival. Both Miksad and colleagues\textsuperscript{23} and Hackshaw and colleagues\textsuperscript{25} showed higher R^2 values between the hazard ratios for progression-free survival and overall survival for trials conducted prior to 1990. A possible explanation for this finding is that increased availability of second-line therapies for patients in more recent trials obscures the relationships between survival and the surrogate endpoints.

Authors of these meta-analyses and others who have commented on these results have generally considered the evidence inadequate to fully support the use of progression endpoints as surrogates for survival in breast cancer studies; the primary concerns are that the correlations are too weak and the predictions are too uncertain.\textsuperscript{22-24,26,27}

**Lung cancer**

Correlations between response rate or TTP and the survival outcome in non-small cell lung cancer (NSCLC) were low, but evidence supports PFS as a surrogate measure in extensive-stage small cell lung cancer (SCLC).\textsuperscript{28} Several published meta-analyses in NSCLC provide predictions for OS treatment effects based on the effects of treatment on the response rate or time to progression, but the correlation values for these analyses are extremely low (Table 1D). Johnson and colleagues\textsuperscript{29} also examined factors that might help explain the relationship between effects on the surrogate endpoint
and survival, such as patient age, performance status, stage of disease, year of trial, and use of rescue (or salvage) treatment, but did not identify any other factor that predicted differences in survival.

A meta-analysis of six single-arm and three randomized trials in patients with extensive-stage SCLC showed correlation between PFS and OS as strong as that seen in colorectal studies. However, the authors of this study cautioned that further validation, using data from a larger number of randomized Phase 3 trials, was needed. The apparently better performance of PFS as a surrogate for OS in SCLC than in NSCLC may relate to the more aggressive untreated clinical course of SCLC compared with NSCLC and the higher responsiveness and greater proportional survival benefit with systemic treatment (at least 4- to 5-fold) for patients with SCLC, compared to those with NSCLC (approximately 33% improvement in median survival, from 4.5 to 6 months).

Renal cell carcinoma
Several approaches have been used to assess the relationship between progression and survival in renal cell carcinoma. Two groups of investigators reviewed results from clinical trials and reported an association between progression-free survival and overall survival, without presenting any analyses to validate progression-free survival as a surrogate endpoint. Kane and colleagues presented their conclusion as support for the approval of sorafenib for the treatment of renal cell carcinoma largely due to the persuasive magnitude of the improvement in progression-free survival.

At the 2009 annual meeting of the American Society of Clinical Oncology, Delea and colleagues presented a meta-analysis of 21 trials of treatments for renal cell carcinoma; they reported strong correlation ($r = 0.69$) between group differences in median time to progression-free survival and overall survival. The results suggested that a 1-month difference in disease progression was associated with a 1.4-month difference in overall survival. This work was cited in other publications and by the Australian Pharmaceutical Benefits Advisory Committee (PBAC) in its decision regarding everolimus. PBAC stated that Delea’s meta-analysis might not be generalizable to the specific context where crossover to the new treatment was prevalent and rapid, and also noted the lack of information in the abstract regarding a surrogate threshold-effect size.

Rather than using meta-analysis, Heng et al used landmark analysis to suggest that PFS at 3 and 6 months can predict survival. However, the effect of early progressors was not factored into these results, since the landmark analysis excluded patients who progressed or died before the landmark timepoint.

Other cancers
In prostate cancer studies, low observed correlations between PFS and OS suggest that PFS is not a useful surrogate for OS in this disease (Table 1E). Progression-free survival has been evaluated as a surrogate for survival in studies of prostate cancer, but the low correlations do not recommend its use for this purpose.

One publication reported strong correlation between event-free survival and OS in head and neck cancer. This meta-analysis provided evidence that event-free survival, defined as time from randomization to locoregional, distant recurrence, or death from any cause, could be used as a surrogate for overall survival in patients with locally advanced head and neck cancer. The analysis was based on a heterogeneous group of clinical trials in which patients were treated with radiotherapy and/or with concomitant, induction, or adjuvant chemotherapy.

One publication reported moderate correlation between 6-month PFS and 1-year OS in patients with glioblastoma multiforme. Recent publications include meta-analyses evaluating PFS as a surrogate for OS in advanced gastric cancer and non-Hodgkin’s lymphoma; a meta-analysis of surrogates in metastatic melanoma came to our attention after the initial literature search.

Discussion
PFS is the most commonly used surrogate measure in studies of advanced solid tumors, but is not universally accepted, and evidence for its validity varies by tumor type. Correlation with OS is reported for a limited number of cancer types, and validation findings vary by the specific cancer indication, patient subgroups, analytical approaches, and the effectiveness of treatments being studied. In summary, published meta-analyses provide good evidence for the use of PFS as a surrogate for OS in advanced colorectal or ovarian cancer. Evidence in other cancers is limited. In breast cancer, for example, a number of papers have examined the topic of surrogate endpoints using meta-analyses, but the variation in findings is not well understood.

A real treatment effect on a progression endpoint may not predict an effect on overall survival for a number of reasons. It remains possible that progression is not a viable surrogate in some cancers, such as prostate cancer. Even when it is suitable as a surrogate, the relationship between treatment
effects on progression and survival are expected to change in newer trials as more-effective treatments emerge or as new mechanisms of action are explored. The issues discussed here contribute to the difficulty of validating progression-related endpoints as surrogates for OS in oncology studies, and partly explain the lack of evidence in some tumor types.

**Study design issues**

Determination of progression endpoints depends on the definition applied, the frequency and methods of monitoring, evaluator objectivity, and the number and location of lesions that are evaluated. For example, if scans to assess progression in a particular trial are performed every 6 months, then the earliest that progression can be documented is at 6 months; whereas if scans are done every 2 months, then progression can be seen earlier. Thus, studies that have different assessment schedules may show very different relationships between progression timing and overall survival, even if the treatment effects on both endpoints were identical.

Several years and large sample sizes are usually required to collect survival data. Patients often take additional anticancer treatments after disease progression; if effective, these subsequent treatments extend survival. Use of subsequent treatments confounds the effects of the initial treatment of interest and introduces variability, making it harder to pick up a survival signal. When large advantages are identified for a new treatment, ethical considerations prompt offering it to all patients, even those in the control arm. Perversely, this means that we may not be able to observe unconfounded overall survival data for treatments that show significant effects on progression. If treatment administered after progression confounds the association, or if survival after progression is relatively long compared with time to progression, then postprogression factors (eg, noncancer deaths, declining sample size, variability in supportive care) may outweigh treatment differences in time to progression. In other cases, the magnitude of improvement in progression-free survival may be insufficient to translate into a survival benefit.

**Clinical and biological issues**

Even when no further anticancer treatment is given after progression, individual patient variation in tumor growth may obscure the relation of tumor response to survival duration. Many traditional cytotoxic chemotherapy agents are mutagenic, and clonal evolution in cancers is a well-documented phenomenon, so it is theoretically possible that clonal evolution could be accelerated by exposure to cytotoxic drugs, which might affect the relation between PFS and survival. Genetic variants in patients and/or tumors can influence treatment response, risk of progression, or survival.

The relationship between progression and overall survival might be more apparent when we are able to select patients who are the most likely to benefit from a particular treatment, such as when a targeted therapy is directed against a mutated protein that occurs in the tumor cells in only a proportion of patients. Until reliable predictive markers and effective targeted therapies are available for most patients with a given type of cancer, we will need to repeatedly reassess the existing evidence in light of new discoveries. Additionally, some targeted therapies, such as those directed at angiogenesis, may induce disease stabilization rather than tumor regression; therefore, tumors may not exhibit the same pattern of disease progression as that seen under treatment with older regimens. It remains to be seen whether endpoints such as progression-free survival can adequately capture the benefits of these drugs for patients.

**Statistical and validation issues**

Even if a treatment extends survival by the same amount as it extends time to progression, the relationship between endpoints could be difficult to portray statistically. For the purposes of evaluating the validity of a surrogate endpoint using meta-analysis, evidence must be available from similar trials that measured survival without confounding; for a new drug or a new indication, there may not be enough previously conducted trials to conduct such an analysis. Meta-analysis results are affected by which studies are included, what endpoints are evaluated, and what analytical methods are used to model the treatment effects. There is no consensus on what level of correlation between hazard ratios is required to consider PFS a useful surrogate for OS. Furthermore, even if the PFS hazard ratio could be used to predict the OS hazard ratio, it would not convey information on how long patients are expected to live. The hazard ratio depicts differences in the probability of the event over time between two groups, but the magnitude of difference in time to event between treatment groups must also be considered when assessing benefit.

Although meta-analysis has been used extensively for showing the relationship between oncology endpoints, a conceptual difficulty arises from the fact that PFS and OS definitions overlap (ie, preprogression time is a subset of OS). This inherent dependency between PFS and OS has typically not been addressed in meta-regressions, but recent methodological developments account for the dependency structure between OS and PFS (eg, by relating
PFS to post-progression survival). Some of these promising approaches include specialized parametric models, simulations, and sophisticated multistate models of joint distribution of endpoints that incorporate the time between progression and survival when making predictions.

**Regulatory issues**

In general, justification for the use of a surrogate endpoint for drug approval depends on such considerations as tumor type, patient population, line of treatment, drug mechanism of action, development phase, and availability of alternate treatments. The European Medicines Agency has accepted PFS and disease-free survival as primary endpoints in situations where there is a “large effect on progression-free survival, a long expected survival after progression, or a clearly favorable safety profile.” In refractory metastatic tumors or when no available alternative therapies exist, the FDA may grant approval based on the effect of treatment on a surrogate endpoint that is “reasonably likely to predict a clinical benefit.”

Draft guidance from the FDA for endpoints in NSCLC states that consideration of PFS as an endpoint for the demonstration of efficacy for drug approval will be based on the magnitude of the effect and the risk-benefit profile of the drug. “Because of the subjectivity in the measurement of PFS assessments and the fact that assessments depend on frequency, accuracy, reproducibility, and completeness, the observed magnitude of effect should be substantial and robust.” In other words, regardless of evidence for the validity of PFS as a surrogate for survival, the magnitude of benefit continues to be a key driver for the approval of new drugs.

**Conclusion**

The interpretation of treatment differences, however expressed, requires some judgment, and the threshold defining a clinically important effect for different tumor types depends on many factors, such as natural history of the disease, size and duration of effect, and available alternative therapy. Given the practice of treatment crossover after disease progression, the growing availability of second- and third-line treatment options, and new drug mechanisms of action, it will become increasingly more difficult to establish the relationship between effects on PFS and OS in additional tumor types using meta-analyses based on previously conducted clinical trials. Thus, the magnitude of a treatment effect and the benefit–risk balance remain important considerations in using progression-related endpoints as surrogate endpoints for survival in oncology research. More methodological advancements that address statistical issues and related clinical interpretation are encouraged.

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