Birth by caesarean section and school performance in Swedish adolescents- a population-based study

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Abstract

Background: Our objective was to assess the impact of obstetric mode of delivery, and in particular birth by Caesarean section (CS), on school performance in adolescents using a large, population-based cohort.

Methods: We extracted data from the Swedish Medical Birth Register and National School Register. We included all live singleton births in Sweden from 1982–1995 (n = 1,489,925). School grades were reported on a scale from 0 to 320, scores less than 160 (i.e. “pass”) were considered to be “poor school performance.” Mode of delivery was categorised as: unassisted vaginal delivery (VD), assisted VD, elective CS and emergency CS. We measured the association between mode of delivery and “poor school performance” using logistic regression. We then used quantile regression to assess the association between mode of delivery and school performance across the distribution of scores. We adjusted for maternal age, parity, small and large for gestational age, gestational age, maternal country of birth, maternal depression, non-affective disorder or bipolar disorder, parental income at time of birth, and parental social welfare at time of birth. We also conducted sensitivity analyses to investigate the association further.

Results: With logistic regression analysis, the adjusted odds ratio (aOR) of assisted VD and poor school performance, compared to unassisted VD, was 1.06 (95% CI: 1.03–1.08). For elective CS it was 1.06 (95% CI: 1.03-1.09) and for emergency CS it was 1.12 (95% CI: 1.09–1.15). With quantile regression, assisted VD showed little difference in scores, when compared to unassisted VD, at any point across the distribution. Elective CS was associated with a 1–3 point decrease in scores, and emergency CS was associated with a 2–5 point decrease in scores.

Conclusion: A slight association was found between birth by CS and school performance. However, the effect was quite small and given the complex nature of the relationship, should be interpreted with caution.

Keywords: Caesarean section: School performance, Cohort study, Quantile regression

Background

Rates of Caesarean section (CS) are rising globally. A recent study including data from thirty one European countries, reported the median rate of CS in those countries to be 25%, [1] though this trend is by no means limited to Europe or other high-income countries. Another recent study reported that in 2010–2011, among the 21 countries included, the rate of CS in countries with a high Human Development Index (HDI) was 40.0%, in countries with moderate HDI it was 32.4% and in countries with low HDI it was 20.3% [2]. Given such a large and growing rate of CS, it is becoming of increasing interest to understand potential long-term effects of birth by CS, [3] as even a small increase in risk could potentially have a large impact globally.

It has been hypothesised that birth by CS leads to changes in psychological development, due to “early” term birth [4, 5] or alterations in microbiota or stress response [6]. In animal models, birth by CS and changes in microbiota have been associated with changes in
behaviour, stress response, and anxiety [7]. Notably, these theories apply more to pre-labour or “elective” CS. Previous evidence suggests the association between CS and neurodevelopmental disorders in human populations may be primarily driven by confounding [8]. However, it is also possible that birth by CS may lead to sub-clinical changes in behavioural development, for example increased anxiety, that are not included in official diagnoses. One way to potentially assess the overall impact on psychological well-being is through school performance. School performance has been associated with several psychological outcomes including behavioural problems, [9] attention problems, [10] substance abuse, [10] and insomnia [11]. In addition, school performance as a teenager has been linked to psychological well-being as an adult, including depression and self-harm [12, 13]. Therefore, if CS has an impact on behavioural development, it may lead to changes in overall school performance.

Elective CS has been associated with delays in personal social skills and gross motor function at age 9 months, [14] and early term birth has been associated with increased special education requirements [4]. However, to our knowledge, no one has examined the potential association between obstetric mode of delivery, more specifically birth by CS, and school performance. To that end, the objective of the current study was to investigate the possible impact of birth by CS on school performance in a large, population-based Swedish cohort.

Methods

Data

Study Population

We included data from 4 Swedish Registers: Medical Birth Register, National School Register, Multi-Generation Register, and National Patient Register. Each resident of Sweden is given a personal identification number (PIN) which is the same in each of these registers, and can be used to link data across registers. The Medical Birth Register was established in 1973 and includes data on over 98% of all births in Sweden [15]. For our cohort, we included all live singleton births in the Swedish Medical Birth Register that occurred between 1982 and 1995. As both mode of delivery and school performance are likely to be highly correlated in multiple births, these were excluded. Variables detailing the timing of onset of labour and CS are available from 1982, thus marking the beginning of our cohort births. Ethical approval was obtained from the regional ethical research committee of Stockholm at Karolinska Institutet. Informed consent was waived by the ethics committee.

Exposure-obstetric mode of delivery

Obstetric mode of delivery, extracted from the Medical Birth Register, consisted of “unassisted vaginal delivery VD,” “assisted VD,” “elective CS” and “emergency CS”. Unassisted VD was defined as VD without the use of forceps or vacuum extraction, and assisted VD was VD with the use of forceps or vacuum extraction. Unassisted and assisted VD included both spontaneous and induced VD. Elective CS was defined as CS which started before onset of labour (as indicated on medical charts by water departure, bleeding or regular and sustained pain) and emergency CS was defined as CS which started after onset of labour.

Outcome-school performance

Data on school performance were extracted from the National School Register, which are available beginning in 1988. In Sweden, upon finishing the compulsory years of school (age 16), grades in 16 subjects are recorded. Starting in 1998, these grades were categorised into 4 levels for each subject: not pass (score of 0), pass (score of 10), pass with distinction (score of 15), and pass with great distinction (score of 20). This allowed for a maximum total score of 320 (i.e. a score of 20 in each of the 16 subjects). Prior to 1998 there was a different grading system, but as the oldest children in our cohort turned 16 in 1998 only the current method was included. Children that “drop out” of school before compulsory grading still technically graduate but are recorded as having received a total of 0 for their final grade, and are not able to continue on to high school. These children were included in our population and were recorded as having a total score of 0. Scores were assessed in both categorical and continuous (from 0 to 320 in jumps of 5) form [16, 17].

Co-variates

Based on previous literature and the use of a directed acyclic graph (DAG) [18] (Additional file 1: Figure S1), the following a priori co-variates were included in the analysis: maternal age at time of birth (<25 years, 25–34 years, 35–44 years, 45+ years), [16] birth order (first born), [16] small for gestational age (SGA), [16] large for gestational age (LGA) (defined as birth weight less or greater than 2 standard deviations from the mean for gestational age, respectively), gestational age (<37 weeks, 37, 38, 39, or 40 weeks, >40 weeks), [16] maternal country of birth (Swedish, other Nordic, other), [19] maternal depression, non-affective disorder, or bipolar disorder (never diagnosed, diagnosed before birth, diagnosed after birth), parental income at time of birth (in quintiles), and parental social welfare at time of birth (yes/no, note: available from 1983), [19] and parental highest education (pre-high school, high school, post-high school) [16, 19, 20].

Though not identified as confounders in the DAG, further co-variates that were identified based on previous literature were also assessed, including: year of birth,
[20] year of school completion, smoking at time of first antenatal visit (none, 1–9 cigarettes/day, 10+ cigarettes/day), [16, 20] infant gender, [16, 19] Apgar score at 5 min (“low” [0–3], “intermediate” [4–6], “high” [7–10]), [16, 21] paternal country of birth (Swedish, other Nordic, other), [19] paternal depression, non-affective disorder, and bipolar disorder (never diagnosed, diagnosed before birth, diagnosed after birth), parental co-habitation at time of birth [19, 20].

All co-variates were tested individually in the logistic regression analysis to assess the potential impact on the association between mode of delivery and school performance. As no variable changed the estimate by more than 10% (only maternal age changed the estimate by more than 5%), only the variables decided on a priori were included in final analysis. Notably, parental education was considered an a priori variable, but was only available from 1990, and thus though it was tested individually, and as it had no impact on the estimate, was not included in the model. Distribution of each variable by mode of delivery is outlined in Table 1.

**Statistical analysis**

**Logistic regression**

For the logistic regression, we considered “poor school performance” to be a total score of less than 160, [16, 17, 22] meaning the individual did not have an average of at least 10 (i.e. “pass”) for the 16 subjects. In Sweden, scores are assigned by teachers rather than a standardised test, and thus standards for a particular grade could vary school-to-school. To account for this, we used mixed effects modelling with a random intercept for school ID.

**Quantile regression**

The data on school performance have been previously reported to be highly skewed [16, 20]. We used quantile regression to analyse school performance in its continuous form. Quantile regression is similar to an ordinary least squares (OLS) model, except the model regresses on the quantile of interest (such as the median), instead of the mean. Quantile regression also does not require an assumption of normality or equal variance, and allows for assessment across the distribution (i.e. at every quantile). In this way we were able to determine if there was an effect of mode of delivery across the distribution of scores (for example, a possible effect only on the high or low scores), rather than an effect only on passing scores as seen with logistic regression. We plotted quantile regression coefficients for every fifth quantile from the 5th to the 95th using the kernel-based method for estimating standard errors [23]. We also looked at coefficient estimates for specifically the 5th, 25th, 50th, 75th, and 95th quantiles. In adjusted analysis we included the same co-variates as the fully adjusted logistic regression model.

**Additional analyses**

We conducted several sub-group analyses. In the logistic regression, we restricted to births from 1990 onwards (the year data on parental education became available), and assessed the association with and without adjustment for parental education. We assessed the association only among male babies. We also excluded children born through a secondary CS (children born by CS whose mothers’ had previously given birth through CS), and children with a low Apgar score at 5 min. Though the vast majority of the population finishes compulsory years of school at age 16 (95%), there are some students who finish younger or older. To that end, we also restricted the population to those who were 16 at the time they finished compulsory school to determine what effect age may have had on school performance. To account for potential clustering of academic performance within families, we restricted the population to one-child families and first born children. We then repeated overall analysis with a random intercept for maternal ID instead of school ID. For both logistic and quantile regression we conducted sensitivity analyses by excluding children who received a “0” as a grade (i.e. children who did not complete the compulsory years of schooling).

Additionally, we conducted logistic regression to assess the association between birth by CS and school performance in five subject categories: [12] natural sciences (biology, chemistry and physics), social sciences (geography, religion, history and society knowledge), arts (art and handicraft), sports, and Swedish. An average below “pass” (10 points per subject) was considered poor performance in each category. Similar to overall school performance analysis, sub-group analyses were conducted where children who were recorded as a “0” in any subject were excluded from that category.

The logistic regression analysis was conducted in SAS v9.3 (Cary, N.C) using PROC GLIMMIX [24] and quantile regression analysis was conducted in R v3.2.2 using the QUANTREG package [23]. Missing data were addressed using the missing indicator method, with a category for each variable used to indicate “missing” status [25].

**Results**

**Descriptive statistics**

There were 1,489,925 live births in Sweden from 1982–1995. We excluded 34,199 (2.3%) multiple births, and 52,562 (3.5%) births missing mode of delivery, leaving a total of 1,403,164 births. Of these, 176,836 (12.6%) did not have grades recorded in the National School Register. Possible reasons for not being recorded in the National School Register include death, emigration, and attendance...
### Table 1 Distribution of descriptive variables by mode of delivery

| Variable                             | Unassisted VD | Assisted VD | Elective CS | Emergency CS |
|--------------------------------------|---------------|-------------|-------------|--------------|
| Total Population                     | 1036424 (84.51) | 78441 (6.40) | 52107 (4.25) | 59357 (4.84) |
| Poor grade                           | 138202 (13.33) | 8880 (11.32) | 7064 (13.56) | 8203 (13.82) |
| Gender (Male)                        | 522988 (50.46) | 45361 (57.83) | 26165 (50.21) | 32357 (54.51) |
| Maternal age                         |               |             |             |              |
| <20                                  | 30367 (2.93)  | 2679 (3.42)  | 698 (1.34)  | 1551 (2.61)  |
| 21–29                                | 633991 (61.17) | 50482 (64.36) | 24004 (46.07) | 33689 (56.76) |
| 31–39                                | 356370 (34.38) | 24176 (30.82) | 2486 (47.71) | 22406 (37.75) |
| 40+                                  | 15696 (1.51)  | 1104 (1.41)  | 2545 (4.88) | 1711 (2.88)  |
| First born child                     | 599463 (57.84) | 69919 (89.14) | 31350 (60.16) | 44922 (75.68) |
| Smoking at time of first antenatal visit |          |             |             |              |
| None                                 | 704557 (67.98) | 52772 (67.28) | 35328 (67.80) | 39483 (66.52) |
| 1–9 Cigarettes/day                   | 148178 (14.30) | 11924 (15.20) | 7746 (14.87) | 9586 (16.15)  |
| 10+ Cigarettes/day                   | 91075 (8.79)  | 6401 (8.16)  | 4979 (9.56)  | 5864 (9.88)  |
| Missing                              | 92614 (8.79)  | 7344 (9.36)  | 4054 (7.78)  | 4424 (7.45)  |
| SGA                                  | 20383 (1.97)  | 2178 (2.78)  | 2969 (5.75)  | 1507 (2.45)  |
| LGA                                  | 31455 (3.03)  | 2279 (2.91)  | 2996 (5.75)  | 2605 (4.39)  |
| Parental education                   |               |             |             |              |
| Pre-Highschool                       | 34151 (3.30)  | 209 (2.67)   | 1878 (3.60)  | 2342 (3.95)  |
| Highschool                           | 232551 (22.44) | 17107 (21.81) | 11862 (22.76) | 15568 (26.23) |
| Post-Highschool                      | 151078 (14.58) | 12916 (16.47) | 8555 (16.42) | 10506 (17.70) |
| Missing                              | 9145 (0.88)   | 680 (0.87)   | 344 (0.66)   | 703 (1.18)   |
| Born Prior to 1990                   | 609499 (58.81) | 45647 (58.19) | 29468 (56.55) | 30238 (50.94) |
| Parental income                      |               |             |             |              |
| First                                | 201292 (19.42) | 8839 (11.27)  | 8943 (17.16) | 9095 (15.32) |
| Second                               | 210046 (20.27) | 9836 (12.54)  | 10020 (19.23) | 9407 (15.85) |
| Third                                | 206797 (20.23) | 13499 (17.21) | 10215 (19.60) | 10713 (18.05) |
| Fourth                               | 204240 (19.71) | 19267 (24.56) | 10316 (19.80) | 13159 (22.17) |
| Fifth                                 | 191366 (18.46) | 25526 (32.54) | 11730 (22.51) | 15816 (26.65) |
| Missing                              | 1981 (1.91)   | 1474 (1.88)   | 883 (1.69)   | 1167 (1.97)  |
| Parental social welfare status       |               |             |             |              |
| Yes                                  | 84609 (8.16)  | 5597 (7.14)  | 4173 (8.10)  | 5490 (9.25)  |
| No                                   | 864190 (83.38) | 65913 (83.03) | 44174 (84.78) | 50000 (84.24) |
| Missing                              | 19132 (1.85)  | 1416 (1.81)  | 838 (1.65)   | 1133 (1.91)  |
| Born Prior to 1983                   | 668493 (66.61) | 5157 (7.03)   | 2902 (5.57)  | 2734 (4.61)  |
| Apgar Score (5 min)                  |               |             |             |              |
| Low (0–3)                            | 1484 (0.14)   | 241 (0.31)   | 141 (0.27)   | 417 (0.70)   |
| Intermediate (4–6)                   | 3375 (0.33)   | 1501 (1.91)  | 479 (0.92)   | 1836 (3.09)  |
| High (7–10)                          | 1004980 (96.97) | 75464 (96.20) | 50105 (96.16) | 55775 (93.97) |
| Missing                              | 26585 (2.57)  | 1235 (1.57)  | 1382 (2.65)  | 1329 (2.24)  |
| Gestational age                      |               |             |             |              |
| <37 Weeks                            | 39384 (3.80)  | 1971 (2.51)  | 5941 (11.40) | 10820 (18.23) |
| 37 Weeks                             | 48430 (4.67)  | 2546 (3.25)  | 5923 (11.37) | 4508 (7.59)  |
| 38 Weeks                             | 122702 (11.84) | 6707 (8.55)  | 25006 (47.99) | 7487 (12.61) |
| 39 Weeks                             | 255027 (24.61) | 15235 (19.42) | 9990 (19.17) | 8416 (14.18) |
| 40 Weeks                             | 311130 (30.02) | 23094 (29.44) | 2714 (5.21)  | 11056 (18.63) |
| >40 Weeks                            | 257116 (24.81) | 28678 (36.56) | 2400 (4.61)  | 16892 (28.46) |
### Table 1 Distribution of descriptive variables by mode of delivery (Continued)

|                        | VD   | CS   | SGA  | LGA  |
|------------------------|------|------|------|------|
| Missing                | 2635 (0.25) | 210 (0.27) | 133 (0.26) | 178 (0.30) |
| Maternal Depression    |      |      |      |      |
| Never Diagnosed        | 952661 (91.92) | 71714 (91.42) | 46936 (90.08) | 53615 (90.33) |
| Diagnosed before birth | 6051 (0.58) | 457 (0.58) | 584 (1.12) | 540 (0.91) |
| Diagnosed after birth  | 77712 (7.50) | 6270 (7.99) | 4587 (8.80) | 5202 (8.76) |
| Maternal Non-affective Disorder |      |      |      |      |
| Never Diagnosed        | 1026908 (99.08) | 77638 (99.98) | 51460 (98.76) | 58644 (98.82) |
| Diagnosed before birth | 2392 (0.23) | 227 (0.29) | 209 (0.40) | 201 (0.34) |
| Diagnosed after birth  | 7124 (0.69) | 576 (0.73) | 438 (0.84) | 501 (0.84) |
| Maternal Bipolar Disorder |      |      |      |      |
| Never Diagnosed        | 1026027 (99.00) | 77649 (99.99) | 51507 (99.23) | 58023 (99.00) |
| Diagnosed before birth | 818 (0.08) | 82 (0.10) | 80 (0.15) | 61 (0.10) |
| Diagnosed after birth  | 9579 (0.92) | 710 (0.91) | 520 (1.00) | 594 (1.00) |
| Paternal Depression    |      |      |      |      |
| Never Diagnosed        | 983156 (94.86) | 74367 (94.81) | 49255 (94.53) | 56114 (94.54) |
| Diagnosed before birth | 4697 (0.45) | 309 (0.39) | 309 (0.59) | 325 (0.55) |
| Diagnosed after birth  | 48571 (4.69) | 3765 (4.80) | 2543 (4.88) | 2918 (4.92) |
| Paternal Non-affective Disorder |      |      |      |      |
| Never Diagnosed        | 1028187 (99.21) | 77816 (99.20) | 51705 (99.23) | 58823 (99.10) |
| Diagnosed before birth | 2360 (0.23) | 171 (0.22) | 134 (0.26) | 167 (0.28) |
| Diagnosed after birth  | 5877 (0.57) | 454 (0.58) | 268 (0.51) | 367 (0.62) |
| Paternal Bipolar Disorder |      |      |      |      |
| Never Diagnosed        | 1029320 (99.31) | 77915 (99.33) | 51703 (99.22) | 58919 (99.26) |
| Diagnosed before birth | 805 (0.08) | 65 (0.08) | 68 (0.13) | 43 (0.07) |
| Diagnosed after birth  | 6299 (0.61) | 461 (0.59) | 336 (0.63) | 395 (0.67) |
| Co-habitating with child’s father |      |      |      |      |
| Yes                    | 894777 (86.33) | 66362 (84.60) | 45489 (87.30) | 50855 (85.68) |
| No                     | 47235 (4.56) | 4655 (5.93) | 2389 (4.58) | 3464 (5.84) |
| Missing                | 94412 (9.11) | 7424 (9.46) | 4229 (8.12) | 5038 (8.49) |
| Maternal Country of Birth |      |      |      |      |
| Sweden                 | 917862 (88.56) | 69758 (88.93) | 45638 (87.59) | 51349 (86.51) |
| Other Nordic           | 41385 (3.99) | 2935 (3.74) | 2327 (4.47) | 2526 (4.26) |
| Other                  | 77170 (7.45) | 5747 (7.33) | 4141 (7.95) | 5482 (9.24) |
| Missing                | 7 (0.00) | 1 (0.00) | 1 (0.00) | 0 (0.00) |
| Paternal Country of Birth |      |      |      |      |
| Sweden                 | 905824 (87.40) | 68936 (87.88) | 45771 (87.84) | 51084 (86.06) |
| Other Nordic           | 34953 (3.37) | 2445 (3.12) | 1910 (3.67) | 2153 (3.63) |
| Other                  | 90567 (8.74) | 6485 (8.27) | 4155 (7.97) | 5713 (9.62) |
| Missing                | 5080 (0.49) | 575 (0.70) | 271 (0.52) | 407 (0.69) |
| Natural science poor performance |      |      |      |      |
| Sweden                 | 298903 (28.43) | 21026 (26.51) | 15266 (28.86) | 17513 (29.09) |
| Other Nordic           | 408082 (38.81) | 49818 (37.18) | 32022 (39.47) | 36182 (39.91) |
| Other                  | 70358 (6.79) | 4365 (5.56) | 3611 (6.93) | 4048 (6.82) |
| Missing                | 62182 (5.91) | 3937 (4.96) | 3223 (6.09) | 3644 (6.05) |
| Swedish poor performance |      |      |      |      |
| Sweden                 | 68154 (6.48) | 4196 (5.29) | 3467 (6.55) | 4297 (7.14) |

*Abbreviations: VD vaginal delivery, CS Caesarean section, SGA small for gestational age, LGA large for gestational age*
at a specialised school (for example if the child had a disability) [21]. There were 143,740 missing that were born through unassisted VD (12.2%), 11,538 from assisted VD (12.8%), 9,615 from elective CS (15.6%) and 11,943 from emergency CS (16.8%). Our final cohort consisted of 1,243,876 children. Of these, 1,036,424 were born by unassisted VD (84.5%), 78,441 by assisted VD (6.4%), 42,107 by elective CS (4.3%) and 59,357 by emergency CS (4.8%) (Table 1). 162,349 (13.2%) children had poor school performance. Ages at time of grading ranged from 14–21, with 95% of children receiving grades at 16 years of age. The median score was 210 (Interquartile range: 175–250) (Fig. 1).

**Logistic regression**

In unadjusted analysis, the OR for assisted VD compared to unassisted VD was 0.84 [95% CI: 0.82–0.86] (Table 2). After adjustment it was 1.06 [95% CI: 1.03–1.08]. For elective CS the unadjusted OR was 1.05 [95% CI: 1.03–1.08] and the adjusted OR was 1.06 [95% CI: 1.03–1.09]. For emergency CS the unadjusted OR was 1.05 [95% CI: 1.02–1.07] and increased to 1.12 [95% CI: 1.09–1.15] after adjustment. Adjusting for birth year, year of receiving grades, and parental education (for those born after 1990) had no effect on results (Additional file 2: Table S1 and Additional file 3: Table S2). Results were consistent among male babies, children born in Stockholm County, children who were 16 years old when they received grades, children without a low Apgar score, and children with a score above 0 (Additional file 4: Table S3). Similarly, there seemed to be no family effect seen by creating a random intercept for maternal ID, or restricting the population based on family size and birth order (Data not shown, available on request).

When school performance was divided into subjects, the association was similar to the overall association between birth by CS and poor school performance (Additional file 5: Table S4).

**Quantile regression**

There was little difference in the distribution of grades by mode of delivery (Fig. 1). The plots of estimates across the distribution, as well an explanation for how to interpret them, is shown in Fig. 2. Scores for children born through assisted VD were slightly higher than unassisted VD, especially at lower quantiles. There was no association with elective or emergency CS. After adjustment, there was no association between assisted VD and school performance (Fig. 3 and Additional file 6: Figure S2). Elective and emergency CS were associated with a slight decrease (2–5 points) in scores across the distribution.

In unadjusted analysis, the 5th, 25th, 50th, 75th and 95th percentiles for unassisted VD were 70, 175, 210, 250 and 300, respectively. Assisted VD was 20 points lower than unassisted VD at the 5th percentile, and 50 points lower at the 95th percentile.
higher for the 5th percentile, and 5 points higher for the 50th and 75th percentiles. Elective CS was the same as unassisted VD at all five percentiles. Emergency CS was 5 points lower than unassisted VD at the 25th and 50th percentiles. In adjusted analysis, the 5th, 25th, 50th, 75th, and 95th percentiles for unassisted VD were 17.2, 132.8, 180, 215, and 274.7, respectively. Assisted VD was largely equal to unassisted VD. Elective CS showed a slight decrease (ranging from $-1.4$ to $-2.78$) in scores across the distribution, with the exception of the 75th percentile where it was equal. Similarly, emergency CS also showed a slight decrease in scores (ranging from $-2$ to $-5$) with the exception of the 75th percentile, which showed no change. Excluding scores of 0 did not change results (Data not shown, available on request).

**Discussion**

**Main findings**

We assessed the impact of obstetric mode of delivery, and in particular birth by CS, on school performance at age 16 using a large, population-based cohort. Two separate analyses were conducted, logistic and quantile regression, assessing school performance in both dichotomous and continuous form. There was a slight association between birth by CS and a reduction in school performance in both analyses. In logistic regression, elective CS was significantly associated with increased likelihood of poor school performance. However, children born by CS were only 6% more likely to receive a poor grade when compared to children born through unassisted VD. The association between emergency CS and poor school performance was somewhat stronger but still small (OR = 1.12, [95% CI:1.09–1.15]). With quantile regression analysis, there again appeared to be a slight association between birth by CS and school performance, primarily in adjusted results. Children born by elective or emergency CS had a 1–5 point decrease in score across the distribution, translating to a 0.31–1.56% decrease.

**Table 2** Association between mode of delivery and poor school performance

|                     | Exposed Cases | Unadjusted OR (95% CI)     | Adjusted OR (95% CI)   |
|---------------------|---------------|----------------------------|------------------------|
| Unassisted VD       | 138202        | Ref                        | Ref                    |
| Assisted VD         | 8880          | 0.84 (0.82-0.86)           | 1.06 (1.03-1.08)       |
| Elective CS         | 7064          | 1.05 (1.03-1.08)           | 1.06 (1.03-1.09)       |
| Emergency CS        | 8203          | 1.05 (1.02-1.07)           | 1.12 (1.09-1.15)       |

Abbreviations: OR odds ratio, VD vaginal delivery, CS Caesarean section

Fig. 2 Unadjusted quantile regression modelling the association between mode of delivery and school performance. *Interpretation: the coefficient values for unassisted vaginal delivery correspond to the school grade at that percentile. Estimated coefficients for other modes of delivery correspond to the estimated difference in score at each percentile. Grey area indicates 95% confidence intervals. Note: Percentiles are recorded as proportions. For example, 0.2 corresponds to the 20th percentile etc.*
Interpretation

There are several possible explanations for an observed association between birth by CS and a small reduction in school performance. A range of characteristics influence school performance including not only behaviour [9] and personality, [26] but also cognitive ability, [26] and external factors such as ethnic diversity in the district [19]. It is possible that rather than having an effect across this wide range of factors and behaviours, birth by CS is having an effect on only one aspect, such as anxiety. Another potential explanation is that this result is being driven by confounding, such as confounding by indication or residual confounding. Confounding by indication occurs when an outcome is causally associated with an indication for the exposure of interest [27]. For example, foetal distress and maternal anxiety may be indications for emergency or elective CS, [28] and may also have an impact on school performance [29, 30], leading to a non-causal association between CS and poor school performance. Additionally, the association could be driven by residual confounding [27]. The relationship between pre and perinatal risk factors and psychological development is complex, and it can be difficult to rule out the effect of difficult to measure confounders, such as social adversity [31]. Regardless of what is driving the association, the decrease in score is very slight. A previous study on this population reported comparable effect sizes for the association between current asthma, rhinitis, eczema and school performance (change in mean score ranging from −3.1 to 4.1) and similarly concluded that though there were statistically significant associations they were likely not clinically meaningful or causally associated [32]. For comparison, the same study reported that severe nasal symptoms is associated with a 12.1 point decrease in mean grade, [32] and another study reported consumption of fish at least once a week is associated with a 14.5–19.9 point increase in mean grade [33].

Strengths and limitations

The strengths of the present study include the use of population-based registries, limiting selection, information bias and measurement error, as we were able to include the entire population and received data from official records. Also, due to the extensive nature of these registries, we were able to assess the effect of a wide variety of co-variates and potential confounders including not only obstetric information, but also demographic and socio-economic factors. Additionally, due to the grading system in Sweden and the use of quantile regression, we were able to assess the impact of mode of delivery across a range of school grades, rather than merely assessing the likelihood of a “passing” grade.

The present study also has several limitations. First, we had no data on breast feeding, which has been linked to both mode of delivery [34] and school performance [35]. However, Sweden has a very high rate of breast feeding and close to 100% of Swedish-born children have ever been breast fed [36]. Additionally, as breast feeding may be affected by mode of delivery it is more likely to
be a mediator rather than a confounder in this situation. Second, a range of factors affect school performance, and we cannot rule out a potential effect on more specific outcomes, such as anxiety, disruptive behaviour, or cognition. Previous results would indicate that mode of delivery does not have an impact on childhood neurodevelopment, [8] but results on behavioural difficulties are conflicting [14, 37]. Finally, it is worth noting that birth by CS in Sweden may not be representative. Access to medical care in Sweden is egalitarian, and the associations between social class and CS seen in other countries are not as prevalent [38]. Additionally, Sweden has a very low rate of birth by CS compared to other European countries, [1] and it is probable that we had a low incidence of non-medically indicated CS. Though we did not have access to information on indications for CS, we were able to separate pre- and post-labour CS, which had no impact on results.

Conclusions
The present study used two analysis methods, adjusted for a wide variety of potential confounders, and conducted several sensitivity analyses to further investigate a potential association. With these robust analysis methods, we have concluded there is a slight association between birth by CS and poor school performance. Given the complex nature of the relationship between perinatal risk factors, such as birth by CS, and development, this small association should be interpreted with caution.

Additional files

Additional file 1: Figure S1. Proposed directed acyclic graph (DAG) describing the association between birth by Caesarean section and poor school performance. Dark grey corresponds to variables which were measured, light grey corresponds to variables which were not measured or are difficult to fully quantify. (JPG 340 kb)

Additional file 2: Table S1. Effect of birth year and year of grading on the association between mode of delivery and poor school performance. (DOCX 15 kb)

Additional file 3: Table S2. Effect of parental education on the association between mode of delivery and poor school performance among children born in 1990 or later. (DOCX 14 kb)

Additional file 4: Table S3. Sensitivity analyses examining the effect of gender, county of birth, age at grading and Apgar score on the association between mode of delivery and poor school performance. (DOCX 15 kb)

Additional file 5: Table S4. The association between mode of delivery and poor school performance by subject. (DOCX 17 kb)

Additional file 6: Figure S2. Complete model of the adjusted association between mode of delivery and school performance. (TIFF 1994 kb)

Abbreviations
aOR: Adjusted odds ratio; CS: Cesarean section; DAG: Directed acyclic graph; HDI: Human development index; LGA: Large for gestational age; PIN: Personal identification number; SGA: Small for gestational age; VD: Vaginal delivery

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Availability of data and materials
The data that support the findings of this study are available from Karolinska Institutet but restrictions apply to the availability of these data, which were used under license for the current study, and so are not publicly available.

Authors’ contributions
EAC contributed to study design, analysis and interpretation, and drafted the initial manuscript. LCK contributed to study design and interpretation, and critical revision of the manuscript. CD contributed to study design and interpretation, and critical revision of the manuscript. PMK contributed to study design and interpretation, and critical revision of the manuscript. JFC contributed to interpretation of the results and critical revision of the manuscript. TGD contributed to interpretation of the results and critical revision of the manuscript. ASK contributed to study conception, design and interpretation, and critical revision of the manuscript. All authors approved the final manuscript as submitted and agree to be accountable for all aspects of the work in ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved.

Competing interests
The authors declare that they have no competing interests.

Consent for publication
Consent for publication was not required, as data were anonymised and informed consent was waived by the ethics committee.

Ethics approval and consent to participate
Ethical approval was obtained from the regional ethical research committee of Stockholm at Karolinska Institutet. Informed consent was waived by the ethics committee.

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References

1. Macfarlane AJ, Blondel B, Mohangoo AD, Cuttin M, Nijhuis J, Novak Z, et al. Wide differences in mode of delivery within Europe: risk stratified analyses of aggregated routine data from the Euro-Peristat study. BJOG. 2016;123(4):559–68.

2. Vogel JP, Betrán AP, Vindevoghel N, Souza JP, Torloni MR, Zhang J, et al. Use of the Robson classification to assess caesarean section trends in 21 countries: a secondary analysis of two WHO multicountry surveys. Lancet Glob Health. 2015;3(5):e260–e70.

3. Blustein J, Liu J. Time to consider the risks of caesarean delivery for long term child health. BMJ. 2015;350:h3410.

4. MacKay DF, Smith GC, Dobbie R, Pell JP. Gestational age at delivery and special educational need: retrospective cohort study of 407,503 schoolchildren. PLoS Med. 2010;7(6):e1000289.

5. Quigley MA, Poulos G, Boyle E, Wolke D, Field D, Alfirevic Z, et al. Early term and late preterm birth are associated with poorer school performance at age 5 years: a cohort study. Arch Dis Child Fetal Neonatal Ed. 2012;97(3):F167–73.

6. Cho CE, Norman M. Cesarean section and development of the immune system in the offspring. Am J Obstet Gynecol. 2013;208(4):249–54.

7. Cryan JF, Dinan TG. Mind-altering microorganisms: the impact of the gut microbiota on brain and behaviour. Nat Rev Neurosci. 2012;13(10):701–12.

8. Curran EA, Dallman C, Kearney P, Kenny LC, Cryan JF, Dinan T, et al. The association between obstetric mode of delivery and autism spectrum disorder: a population-based sibling design study. JAMA Psychiatry. 2015;72(9):935–42.

9. Poropat AE. A meta-analysis of adult-rated child personality and academic performance in primary education. Brit J Educ Psychol. 2014;84(2):239–52.

10. McLeod JD, Uemura R, Rohman S. Adolescent mental health, behavior problems, and academic achievement. J Health Soc Behav. 2012;53(4):482–97.

11. Kronholm E, Pusniaukka R, Jokela J, Willberg J, Uutila AS, Paunio T, et al. Trends in self-reported sleep problems, tiredness and related school performance among Finnish adolescents from 1984 to 2011. J Sleep Res. 2015;24(1):1–10.

12. Jablonska B, Lindberg L, Lindblad F, Rasmussen F, Ostberg V, Hijen A. School performance and hospital admissions due to self-inflicted injury in adolescence. Pediatr Surg Int. 1999;14(2):130–3.

13. McCarty CA, Mason WA, Kosterman R, Hawkins JD, Lengua LJ, McCauley E. Adolescent school failure predicts later depression among girls. J Adolesc Health. 2008;43(2):180–7.

14. Khalaf SY, O'Neill SM, O'Gorman TD, van Leeuwen WA, D'Onofrio BM. Use of the Robson classification to assess caesarean section trends in 21 countries: a secondary analysis of two WHO multicountry surveys. Lancet Glob Health. 2015;3(5):e260–e70.

15. Stuart A, Otterblad Olausson P, Källen K. Apgar scores at 5 minutes after birth in relation to school performance at 16 years of age. Obstet Gynecol. 2011;118(2, Part 1):201–8.

16. D'Onofrio BM, Richert ME, Frans E, Kuja-Halkola R, Almqvist C, Spolander A, et al. Maternal age at childbearing and offspring psychiatric and academic morbidity. JAMA Psychiatry. 2014;71(4):432–8.

17. Koenker R. Quantile regression in R: a vignette. 2015 [cited; Available from: http://cran.r-project.org/web/packages/quantreg/vignettes/qreg.pdf.

18. U J, Altecor T, Deddings JA. Analysis of large hierarchical data with multilevel logistic modeling using PROC GLIMMIX. 2016 [cited 2018 May 18] Available from: http://www2.sas.com/proceedings/sugi51/151-31.pdf.

19. Rosander P, Bäckström M, Stenberg G. Personality traits and general intelligence as predictors of academic performance: A structural equation modelling approach. Learn Individ Differ. 2011;21(5):590–6.

20. Bosco JL, Silliman RA, Thwin SS, Geiger AM, Buist DS, Proust MN, et al. A most stubborn bias: no adjustment method fully resolves confounding by indication in observational studies. J Clin Epidemiol. 2010;63(1):64–74.

21. Stjernholm YV, Petersson K, Eneroth E. Changed indications for cesarean section. Acta Obstet Gynecol Scand. 2010;89(1):49–53.

22. Galler JR, Ramsey FC, Harrison RH, Taylor J, Cumberbatch G, Forde V. Postpartum maternal moods and infant size predict performance on a national high school entrance examination. J Child Psychol Psychiatr. 2004;45(6):1064–75.

23. van Handel M, Swaab H, de Vries LS, Jongmans MJ. Long-term cognitive and behavioral consequences of neonatal encephalopathy following perinatal asphyxia: a review. Eur J Pediatr. 2007;166(7):645–54.

24. Thapar A, Rutter M. Do prenatal risk factors cause psychiatric disorder? Be wary of causal claims. Br J Psychiatry. 2009;195(2):100–1.

25. Sundberg R, Toren K, Hoglund D, Aberg N, Brismar J. Nasal symptoms are associated with school performance in adolescents. J Adolesc Health. 2007;41(4):581–3.

26. Kim JL, Winkvist A, Aberg MA, Aberg N, Sundberg R, Toren K, et al. Fish consumption and school grades in Swedish adolescents: a study of the large general population. Acta Paediatr. 2010;99(1):72–7.

27. Prior E, Santhakumar S, Gale C, Philips LH, Modi N, Hyde MJ. Breastfeeding after cesarean delivery: a systematic review and meta-analysis of world literature. Am J Clin Nutr. 2012;95(5):1113–35.

28. Heikkila K, Kelly Y, Renfrew MJ, Sacker A, Quigley MA. Breastfeeding and educational achievement at age 5. Matern Child Nutr. 2010;6(3):420–30.

29. OECD Family Database. CO1.5: Breastfeeding rates. 2009. [cited 2016 May 8] Available from: http://www.oecd.org/els/family/43136964.pdf.

30. Thapar A, Rutter M. Do prenatal risk factors cause psychiatric disorder? Be wary of causal claims. Br J Psychiatry. 2009;195(2):100–1.

31. Ducheneau ND, Toffle E, Steelffd E. Obstetric mode of delivery and childhood behavior and psychological development in a British Cohort. J Autism Dev Disord. 2015;45(6):163–14.

32. Thomas J, Hildingsson I, Sweden. In: Kennedy P, Kodeate N, editors. Maternity Services and Policy in an International Context: Risk, Citizenship and Welfare Regimes: Routledge; 2009.