Predictive value of indicators for identifying child maltreatment and intimate partner violence in coded electronic health records: a systematic review and meta-analysis

Shabeer Syed 1,2, Rachel Ashwick 2, Marco Schlosser 3
Arturo Gonzalez-Izquierdo 4, Leah Li 1, Ruth Gilbert 1,4

ABSTRACT
Objective Electronic health records (EHRs) are routinely used to identify family violence, yet reliable evidence of their validity remains limited. We conducted a systematic review and meta-analysis to evaluate the positive predictive values (PPVs) of coded indicators in EHRs for identifying intimate partner violence (IPV) and child maltreatment (CM), including prenatal neglect.

Methods We searched 18 electronic databases between January 1980 and May 2020 for studies comparing any coded indicator of IPV or CM including prenatal neglect defined as neonatal abstinence syndrome (NAS) or fetal alcohol syndrome (FAS) against an independent reference standard. We pooled PPVs for each indicator using random effects meta-analyses.

Results We included 88 studies (3875183 individuals) involving 15 indicators for identifying CM in the prenatal period and childhood (0–18 years) and five indicators for IPV among women of reproductive age (12–50 years). Based on the International Classification of Disease system, the pooled PPV was over 80% for NAS (16 studies) but lower for FAS (<40%; seven studies). For young children, primary diagnoses of CM, specific injury presentations (eg, rib fractures and retinal haemorrhages) and assaults showed a high PPV for CM (pooled PPVs: 55.9%–87.8%). Indicators of IPV in women had a high PPV, with primary diagnoses correctly identifying IPV in >85% of cases.

Conclusions Coded indicators in EHRs have a high likelihood of correctly classifying types of CM and IPV across the life course, providing a useful tool for assessment, support and monitoring of high-risk groups in health services and research.

INTRODUCTION
Intimate partner violence (IPV) and child maltreatment (CM) are forms of family violence that often go unnoticed by services,1–3 despite repeated recommendations by the WHO to improve monitoring efforts.4–6 CM and IPV refer to any act of commission or omission that causes biopsychosocial harm to a child, a future child or partner.6,7 Statutory definitions of CM in the UK include fetal alcohol syndrome (FAS) and neonatal abstinence syndrome (NAS) due to neglect or harm during pregnancy.8 Assessing health records for detailed information on family violence is time consuming and expensive. Instead, studies and services are increasingly using routinely coded electronic health records (EHRs) for assessing family violence.9 10 Coded EHRs allow for longitudinal population-based assessments, automated early warning systems and identification of high-risk populations for targeted interventions at relatively low costs.11–13 However, the potential utility of EHRs to support surveillance and clinical decisions is often undermined by reported quality issues and coded conditions are rarely validated externally.14 15 Unable to check the data themselves for accuracy, large-scale studies (eg, Global Burden of Disease Study) and services rely on routinely coded indicators based on the International Classification of Diseases (ICD) with unknown predictive values.10 16 17 The validity of coded indicators...
are compounded by varying case definitions, ad hoc classifications by coders\textsuperscript{15} \textsuperscript{18–20} and under-recording due to clinician fears of potential harm or lack of awareness.\textsuperscript{18} \textsuperscript{21}

To our knowledge, no previous review has estimated the positive predictive values (PPVs) of coded indicators for different forms of family violence including CM, prenatal neglect (NAS or FAS) and IPV based on external independent reference standards (ie, not using other codes in the EHR to validate indicators). This meta-analysis provides a comprehensive summary of PPVs for multiple coded indicators in EHRs aimed at identifying family violence in general healthcare settings, compared with an independent reference standard.

METHODS

We followed the Preferred Reporting Items for Systematic Reviews and Meta-analyses for Diagnostic Test Accuracy Studies and the Meta-analysis of Observational Studies in Epidemiology guidelines.\textsuperscript{22} \textsuperscript{23} The review protocol was published in the PROSPERO registry (CRD 42019139300),\textsuperscript{24} with any protocol deviations provided in online supplementary table S1.

Search strategy

We searched 18 electronic databases and 20 selected journals for studies published between 1 January 1970 and 24 May 2020 (complete search strategy in online supplementary table S2). Reference lists of eligible full-text articles were hand-searched and 18 frequently cited authors of eligible studies were contacted for article recommendations.

Table 1 Key definitions of electronic health records (EHRs) and family violence

| EHRs | Routinely collected healthcare data stored and accessible electronically as part of routine care or as part of administrative databases for wider patient, practice and billing management purposes. |
| Coded indicators of EHRs | Single or combination of codes aimed at identifying different forms of family violence (table 2). Codes must be based on a recognised classification system (eg, International Classification of Disease system) assigned to patients’ EHRs by a clinician or by a professional coder based on medical records (eg, discharge documentation). |
| Family violence outcomes | Any of the below outcomes obtained independently from the coded indicator (ie, not using other codes in the EHR to validate the indicator): |
| Child maltreatment (CM) | Any act of commission or omission by a parent or caregiver resulting in harm, the potential for harm or threat of harm to a child (0–12 years) or an adolescent (13–19 years) including neglect, psychological, physical, sexual and emotional abuse. Harm does not need to be intended.\textsuperscript{1} \textsuperscript{18–20} |
| Neonatal abstinence syndrome (NAS) or fetal alcohol syndrome (FAS) | Statutory guidelines definition of CM in the UK include any form of maternal omission or commission during pregnancy with potential for significant biopsychosocial harm to the future child such as NAS and FAS.\textsuperscript{3} This may include compliance with prescribed medication (clinician approved) and/or maternal substance misuse during pregnancy, resulting in drug withdrawal symptoms in newborns and/or harm to the future child’s biopsychosocial functioning/appearance. |
| Intimate partner violence (IPV) | Any incident of threatening behaviour, violence or abuse (psychological, physical, sexual, financial or emotional) between adults who are, or have been, intimate partners or family members.\textsuperscript{1} In this review, studies are predominately restricted to estimates of women aged 15–50 years. |

Indicators and outcomes of family violence

An overview of key definitions of indicators and outcomes is provided in tables 1–2. Briefly, indicators were defined as any coded marker or risk factor for family violence, ranging from specific primary diagnostic codes (eg, ‘T74.1 physical abuse’) to injuries (eg, rib fractures and retinal haemorrhages), assaults and combinations of adversity-related codes (table 2). For IPV, we predominately included indicators for women (eg, >80% of study sample were women), as the health consequences for women are higher, and the prevalence among men is significantly lower.\textsuperscript{25} In terms of family violence, men’s parental status is also more difficult to ascertain than women’s. Outcomes were defined as mutually exclusive categories of family violence according to different life periods: NAS or FAS (representing the prenatal period),\textsuperscript{8} any form of CM (representing childhood) or IPV (representing

Study eligibility

Three independent reviewers (SS, RA and MS) screened abstracts, full texts and corresponding reference lists of articles using Covidence’s systematic review software.\textsuperscript{25} Any disagreement over study inclusions between reviewers was resolved by a family violence expert (RG). We included: (1) studies with data to calculate the PPV of a coded EHR indicator for a specific family violence outcome as reported by at least three other eligible studies,\textsuperscript{26} (2) studies from primary care, paediatric units (including trauma centres) or general hospital settings, (3) studies published in English/Swedish/German and (4) studies distinguishing family violence cases from non-cases (tables 1–2; online supplementary table S3–S4).
### Table 2  Coded indicators of specific family violence outcomes as measured by the included studies

| Outcome                                                                 | Indicator and period* | Descriptions*                                                                 | No. data sources† | Frequently used codes‡ |
|------------------------------------------------------------------------|-----------------------|-------------------------------------------------------------------------------|-------------------|------------------------|
| Prenatal neglect/substance misuse during pregnancy                     |                       |                                                                               |                   |                        |
|                          |                       | Diagnosis of substance withdrawal symptoms in newborns, typically within the first few days after birth, resulting from drug use (commonly opioids) during pregnancy. 116 | 13*6 39-43 45     | ICD-9: 779.5, ICD-10: P96.1, |
|                          | Neonatal abstinence syndrome (NAS) |                       |                                                                               |                   |                        |
|                          |                       | NAS: primary diagnosis                                                                                                         |                   |                        |
|                          | Newborn affected by drugs |                       | Less specific indicator of NAS, defined by the fetus being affected by maternal substances, noxious influences or medications NOS. | 7*84            | ICD-9: 760.72, ICD-10: P04.4x, |
|                          | Fetal alcohol syndrome (FAS) |                       | FAS is one type of a spectrum of disorders linked to alcohol exposure during pregnancy. Symptoms include abnormal facial features, deficiencies in growth (including brain/ neurophysiology) and/or neurobehavioural impairment in children (0–19 years, typically diagnosed 2–5 years). 116 | 7*46-52         | ICD-9: 760.71, ICD-10: Q86.0, P04.3, |
| Child maltreatment (CM) among children aged 0–19 years                |                       |                                                                               |                   |                        |
|                          |                       | CM primary diagnoses                                                                                                           |                   |                        |
|                          | Abusive head trauma (AHT) | Skull and/or intracranial injuries due to inflicted blunt impact/penetration and/or violent shaking (0–5 years). Coding clusters include different degrees of certainty: 'definite or presumptive' and 'probable' AHT. Other common terms are 'shaking infant syndrome' or 'whiplash shaking baby syndrome'. 117 | 4*68 80 138      | ICD-9: As in CM + 781.0-781.4, 800.x-801.x, 854.0.x, 950.0-950.3, ICD-10: As in CM + S06.0-506.9, S08.7-508.9, T90.5, |
|                          | Traumatic brain injury (TBI) | Skull and/or intracranial injuries and concussion (<5 years of age) characterised by the level of loss of consciousness or altered mental status, ranging from mild TBI (in most cases a concussion) to severe TBI (patient is comatose, unable to open their eyes and so on). 118 | 7*66 69 76 90 34 92 | ICD-9: 800.0-801.9, 802.0-804.9, 850.1-850.5, 851.0-854.1, 959.01, ICD-10: T02.0, T02.1x, T02.8, T02.9, S04.02, S04.04x, S06.x, S07.1, T74.4, |
|                          | Skull fractures | A fracture of one or more bones of the cranial vault or skull base.                                                            | 10*53 57 70 72 73 75 83 88 95 96 | ICD-9: 800.x-804.x, ICD-10: S02.x, |
|                          | Subdural haematoma | A collection of blood between the dura and arachnoid coverings of the brain. In this review, subdural haematoma includes ICD-coded subdural haemorrhages. | 5*57 68 88 94 96  | ICD-9: 852.2, 432.1, ICD-10: 162.0, S06.57, |
|                          | Retinal haemorrhage | Extravasation of blood into the retina, between the retina and the subhyaloid or preretinal or between the retina and the subretina.    | 4*61 88 99 96   | ICD-9: 362.81, ICD-10: H35.6, |
|                          | Upper limb fractures | In this review, referred to as fractures of the humerus, radius or the ulna.                                                     | 4*14 68 99 91 96 | ICD-9: 812.x, 813.x, ICD-10: S42.2-S42.9, T10, |
|                          | Lower limb fractures | In this review, referred to as fractures of the femur, diaphysis, tibia or fibula.                                              | 6*56 63 67 89 96 | ICD-9: 820.0, 821.x, ICD-10: 572.x, S92.x, T12, |
|                          | Rib fractures        | A break in the rib bone of the thoracic skeleton.                                                                              | 4*47 74 88 89   | ICD-9: 807.0-807.1, ICD-10: 522.2, S02.4, |
|                          | Multiple burns       | Burns and corrosions of multiple and unspecified body regions.                                                                   | 3*88 96         | ICD-9: 946.0, ICD-10: T23.x-T23.x, |
|                          | Assaults/CAMadversity | Combination of multiple broad codes deemed to meet the threshold for potential child harm including CM, assaults, drug or alcohol intoxication, poisoning of undetermined cause, concerns for psychosocial circumstances, specific fractures and/or sexual diseases in children (0–18 years). 119 | 6*59 77 78 82 88 | ICD-9: E960.0-E969.0, V71.6, ICD-10: X85-Y09, |
|                          | Intimate partner violence (IPV) among women aged 12–55 years |                                                                 | 18              | ICD-9: As in CM + V15.4, V61.2, V71.8, E960.1, E967, ICD-10: X85-Y09, Z04.0-Z04.8, Z60.0-Z61.8, |

Continued
### Table 2

| Outcome | Description* | No. data sources† | Frequently used codes‡ |
|---------|--------------|-------------------|------------------------|
| IPV     | IPV primary diagnosis Any diagnosis of partner maltreatment syndromes including neglect, abandonment, physical, sexual and psychological abuse.140 | 140 | ICD-9: E967.3, 995.80, 995.81. ICD-10: T74.1, Y06.0, Y07.0. |
| Assaults| Intentional injuries (sexual, physical, emotional or psychological) inflicted by another person. 10 | 10 | ICD-9: E960.0-E969.0, V71.6, V71.5. ICD-10: X85-Y09. |
| Head, face and neck injuries | Maxillofacial fractures (mandible fractures, zygomatic complex fractures and intracranial injuries) and crushing injuries of head and neck. | 10 | ICD-9: 802.0-802.9. ICD-10: S02.2, S02.6, S06.x-S07.x. |
| Ocular injuries | Ocular injuries such as fractures to the orbital floor and sustained scleral laceration or ruptures. | 5 | ICD-9: 802.6, 871.0, 871.1. ICD-10: S02.3, S05.1, S05.2-S05.3. |
| Contusions of upper body | Bruises or contusions of the throat, breast, upper arms/shoulders and eyelids, periocular areas and scalp. | 3 | ICD-9: 920.x-924.x. ICD-10: S10.0-S10.9, S20.0, S40.x, S05.1, S05.7. |

*Indicator definitions represent condensed descriptions for the overall category and are not representative of all ICD code descriptions. **Number of studies contributing to the indictor. †Frequently used codes: ICD-9 codes 48-71, ICD-10 codes T74.1-T74.9. X indicates any character from fourth through seven. ICD, International Classification of Diseases; NOS, not otherwise specified.

### Results

#### Study characteristics

Details of the included study characteristics are provided in online supplementary tables S6–S8 and tables S12–15. In total, women of reproductive age, respectively. Outcomes had to be ascertained and verified by an independent reference standard such as recorded multidisciplinary decisions of family violence extracted from chart reviews (table 1).

#### Data extraction and quality assessments

Using a piloted standardised extraction form, three reviewers independently extracted relevant study characteristics. If studies reported separate estimates for multiple codes, reference standards or age criteria, we extracted all estimates and prioritised those based on the criterion most frequently used by other studies to increase overall homogeneity. We requested additional information from 25 authors, 6 of whom responded within the 2-month deadline and were included. The risk of bias was assessed by the same three reviewers using a revised version of the Quality Assessment of Diagnostic Accuracy Studies tool (QUADAS-2; online supplementary table S4).28 We rated each study’s reference standard according to assessment quality and exclusion criteria based on previous reviews (table 1).29

#### Data synthesis

The PPVs for each family violence outcome were calculated as the proportion of identified cases by an indicator validated as true cases by an independent reference standard. For studies that reported only sensitivity and specificity, we obtained PPVs using Bayes’ theorem.30 PPVs were pooled using random effects intercept logistic regression models with the logit transformation when at least three studies were available for the same outcome.31 The model accounts for potentially misleading back transformations of PPVs when pooling studies with highly variable sample sizes.32 Where applicable, we also examined documentation quality and coding errors by pooling the proportions of coded medical charts with missing key information (full procedures in online supplementary table S5).

We measured the extent to which the PPVs varied between studies (ie, between-study heterogeneity).32 As the PPV is a measure of proportions, we measured the between-study heterogeneity using the I 2 statistic (≥75%=indicates substantial heterogeneity),32 standard χ 2 tests, prediction intervals (where the PPV is expected for 95% of similar future studies)32 and subgroup analyses when at least four studies were available for each subgroup.29 We used random effects meta-regression to assess for the impact of publication year on the between-study heterogeneity. The influence of individual studies was explored by serially omitting different studies from the overall estimates.

Publication bias occurs when studies with favourable results (ie, high PPV) are more likely to be published than unfavourable results.33 We explored publication bias by plotting the PPV against the SEs of individual estimates using funnel plots for indicators with at least 10 studies for the same outcome. To test for statistically significant differences in funnel plot asymmetry (ie, publication bias), we used Egger’s test and Begg’s and Mazumdar rank correlation test.33 We used R (V3.6.1) and the ‘meta’ package with the ‘metaprop’ command to perform the analyses.34 35

#### RESULTS

##### Study characteristics

Details of the included study characteristics are provided in online supplementary tables S6–S8 and tables S12–15. In total,
65 cross-sectional and 23 longitudinal studies (81 unique publications), involving 20 indicators and 3,875,183 individuals from 11 different countries met the inclusion criteria (figure 1). Overall, 13 studies provided indicators for NAS, 7 for FAS, 50 for CM (0–18 years) and 18 for IPV among women (12–50 years). Most studies were from the USA (72 studies, 81%), with a minority from Australia (8 studies, 9%) and Europe (4 studies, 4%). The majority of indicators comprised different ICD-9 coding clusters (64 studies) or modified versions of ICD-10 (14 studies). Chart reviews with predefined family violence criteria were the most frequent type of reference standard for all outcomes (75 studies, median sample size: 301 participants). The smallest study included 38 children, involving all ICD-9 coded subdural haematomas (SDH) at one hospital over 10 years. The study was retained as it involved all potential cases presenting to a generic children’s hospital and as SDHs are deemed to be extremely rare in young children and an important indicator for CM. We included one unpublished dissertation from 2019, specifically assessing ICD codes for CM.

### Study quality assessment

Individual QUADAS-2 scores for studies are provided in online supplementary table S7. Overall, 31 studies (35%) were rated as high risk of bias and lower quality in most domains, and 57 studies (65%) were rated as low risk of bias and higher quality in most domains. The majority of studies were rated lower as they...
did not mask the outcome when reviewing coded EHRs (95%) or used a lower rated reference standard (42%; a rating <4 for CM; rating <3 for NAS, FAS or IPV).

### Table 3: Pooled positive predictive values for NAS, FAS, CM and IPV

| Family violence outcome | Indicator and age cut-off in years | Studies | Ref std. cases/indicator positive | Pooled PPV (95% CI) | Between-study heterogeneity |
|-------------------------|------------------------------------|---------|-----------------------------------|---------------------|-----------------------------|
|                         |                                    |         |                                    |                     | (95% prediction interval)* |
| Prenatal                | NAS primary diagnosis              | 16      | 3030/3796                          | 80.9 (71.0 to 87.9) | (32.3–97.4)                |
|                         | Newborn affected by drugs          | 7       | 384/2202                           | 27.2 (16.8 to 40.8) | (6.5–46.7)                |
|                         | FAS primary diagnosis              | 7       | 777/2687                           | 39.3 (25.3 to 55.4) | (9.7–79.6)                |
| Childhood               | CM                                 | 19      | 3090/3717                          | 87.8 (83.4 to 91.2) | (63.5–96.8)                |
|                         | Abusive head trauma ≤3†             | 4       | 437/477                            | 91.6 (88.8 to 93.8) | (83.8–96.7)                |
|                         | Traumatic brain injury ≤2           | 7       | 410/1770                           | 22.9 (15.3 to 32.9) | (7.0–53.9)                |
|                         | Skull fractures ≤4                  | 10      | 603/2872                           | 22.1 (18.5 to 26.2) | (13.2–34.7)                |
|                         | Subdural haematoma/haemorrhage ≤3†  | 5       | 118/211                            | 55.9 (49.2 to 62.5) | (49.2–62.5)                |
|                         | Retinal haemorrhages ≤3†            | 4       | 60/74                              | 81.1 (70.5 to 88.5) | (70.5–88.5)                |
|                         | Upper limb fractures ≤3             | 6       | 90/272                             | 24.0 (12.5 to 41.1) | (4.7–66.8)                |
|                         | Lower limb fractures ≤3             | 4       | 90/126                             | 88.3 (55.2 to 97.9) | (17.7–99.6)                |
|                         | Rib fractures ≤4                    | 3       | 70/261                             | 22.5 (8.4 to 47.7)  | (3.3–70.9)                |
|                         | Multiple burns ≤2                   | 6       | 218/331                            | 77.0 (46.6 to 92.7) | (19.1–97.9)                |
|                         | Assaults ≤19                        | 3       | 1893/4033                          | 55.9 (37.0 to 73.3) | (14.4–90.5)                |
| Women predominantly of reproductive age | IPV primary diagnosis ≤55 | 3       | 1946/2438                          | 86.1 (72.2 to 93.6) | (53.0–97.1)                |
|                         | Ocular injuries ≤50†                | 4       | 22/668                             | 3.3 (2.2 to 5.0)     | (2.2–5.0)                  |
|                         | Facial fractures ≤50                | 4       | 58/1550                            | 11.1 (3.0 to 33.9)  | (0.6–70.5)                 |
|                         | Upper body contusions ≤50           | 3       | 85/325                             | 26.5 (18.1 to 37.0) | (13.7–44.9)                |
|                         | Assaults ≤65                        | 10      | 6073/20919                         | 31.6 (22.3 to 42.7) | (9.2–67.9)                 |

Each pooled positive predictive value represents the overall logit transformed estimate by indicator and outcome. Pooled estimates are based on random effects meta-analyses, unless otherwise specified. Individual forest plots are shown in the supplement.

*Estimates based on a fixed-effects meta-analysis due to lower between-study heterogeneity (ie, $I^2 <50\%$).

CM, child maltreatment; EHR, electronic health record; FAS, fetal alcohol syndrome; IPV, intimate partner violence; NAS, neonatal abstinence syndrome; PPV, positive predictive value; Ref std, reference standard.

---

**Pooled estimates of PPVs**

Table 3 depicts details for each pooled PPV and between-study heterogeneity by indicator and outcome. Individual forest plots of PPVs and study-specific ICD codes are provided in online supplementary figures S1-S19 and tables S6-S8.

### Neonatal abstinence syndrome

The pooled PPV of primary ICD diagnoses for NAS was 80.9% (95% CI 71.0% to 87.9%), and the PPV ranged from 31.8% to 98.2%, with substantial between-study heterogeneity. Most studies (54%) used the Finnegan scale (a validated 21-item scale for documented symptom severity) to determine the accuracy of the coded diagnosis, with a cut-off score of 8 as a reference standard (except for one study, cut-off score of 4). The remaining studies used recorded clinician diagnosis or required pharmacological NAS treatment as reference standards.

### Fetal alcohol syndrome

For FAS, we found a low pooled PPV of 39.3% (95% CI 25.3% to 55.4%), and the PPV ranged from 14.0% to 66.9%, with large between-study heterogeneity. All studies used ICD-9 codes to identify cases (eg, alcohol affecting fetus) and required 3–5 prespecified criteria to be met in chart reviews, including facial

---

**Child maltreatment**

The pooled PPV of primary diagnoses for CM was 87.8% (95% CI 83.4% to 91.2%) among children aged 0–18 years across 19 studies, with significant between-study heterogeneity. Individual PPVs from studies ranged from 65.0% to 100.0%.

Overall, 37 studies assessed 11 different indicators of CM (online supplementary figures S3-S14). The pooled PPVs ranged from 88.3% (95% CI 55.2% to 97.9%) for rib fractures to 19.6% (95% CI 8.9% to 37.9%) for multiple burn injuries in children under 5 years. The between-study heterogeneity ranged from small to large ($I^2$ range 0.0%–98.5%). Four studies also assessed the PPV of poisonings (range 7.0%–95.0%), but the extreme heterogeneity precluded reliable pooling (prediction interval: 0%–100%). The majority of studies assessing injury indicators of CM actively excluded transport injuries (28 studies, 65%), metabolic bone diseases (10 studies) and birth injuries (9 studies) from the CM group.

### Subgroup analyses for NAS, FAS and CM

Details of each subgroup analysis are provided in online supplementary tables S9–S11. For primary diagnoses of CM, we found significantly higher PPVs of CM in studies from inpatient settings (PPV=90.6%) compared with studies from emergency departments (EDs; PPV=80.8%) and in studies that applied a

---

**Hadfield, S, et al. Arch Dis Child 2021;106:44–53. doi:10.1136/archdischild-2020-319027**
lower rated reference standard (89.4%; a rating <4) relative to a higher rated reference standard (70.8%; a rating ≥4). We found no significant differences when comparing subgroups by coding systems (ICD-9 vs ICD-10), age (younger vs older children) or publication year (p>0.179).

Intimate partner violence
Three studies assessed primary ICD diagnoses of IPV among women presenting to EDs (main age range: 12–55 years), with a pooled PPV of 86.1% (95% CI 72.2% to 93.6%). Individual PPVs from studies ranged from 73.6% to 94.4%. Two studies used the Fitlcraft criteria as a reference standard (eg, specified IPV criteria) and included both violence by ex and current partners as cases. One study from Hong Kong included only IPV within cohabiting couples as cases and used medical chart reviews to verify that diagnoses matched any documented disclosure of IPV.

We found 16 studies (10 in EDs/6 inpatient) that assessed four different injury-related presentations of IPV among predominantly women (primary age range: 12–55 years). The pooled PPVs were low, ranging from 31.6% (22.3%–42.7%) for assault-related codes to 3.3% (2.2%–5.0%) for orbital floor fractures. All studies used ICD-9 codes to identify presentations, except for one Finnish study that used ICD-10 codes. The between-study heterogeneity ranged from low to substantial across indicators ($I^2$ range = 0.0%–99.4%). For assaults and facial fractures, meta-regressions showed that more recent studies were associated with a lower PPV of IPV (p<0.001). Three studies involved a small proportion of men (4%–17%), as estimates could not be separated from women. No study focused on pregnant or elderly women.

Coding errors for NAS, CM and IPV
The proportion of misclassifications (false positives) due to coding errors were on average 2.1% (95% CI 0.8% to 5.6%) across nine studies on NAS, CM, IPV and assault (range $I^2$ = 97.3%–99.7%; online supplementary figure S20).

Documentation quality of assault indicators for IPV
We found that on average 28.0% (95% CI 14.7% to 46.8%; 18074/31224 charts; six studies) of assault coded cases among women had no recorded perpetrator information in the underlying medical charts (ie, missing data), preventing any classification/coding of IPV (online supplementary figure S21; table S5).

Sensitivity analyses
Serially excluding individual studies revealed that no study significantly impacted the between-study heterogeneity, and it remained substantial across all pooled PPVs (online supplementary tables S9-S10).

Publication bias
The funnel plot of studies reporting on the PPV for primary CM diagnoses was asymmetric (Egger’s test p=0.001, slope=0.999, rank correlation test: p=0.248; online supplementary figure S23), meaning that studies with higher PPVs were potentially more likely to be published. We found no evidence of funnel plot asymmetry for all other indicators with at least 10 or more studies (Egger’s test: p>0.366, rank correlation test: p>0.190; online supplementary figures S22-25).

DISCUSSION
This is the largest meta-analysis to investigate the predictive value of indicators for family violence in EHRs, involving over 3.8 million individuals across 11 countries. Despite the large between-study heterogeneity, the results highlight that EHRs provide consistently high PPV for CM or IPV. We found that more than 8 in 10 coded primary diagnoses were confirmed as cases of NAS, CM and IPV. The findings also indicate that 8 in 10 recorded rib fractures and retinal haemorrhages met criteria for CM, and 1 in 3 assault-related presentations among women met criteria for IPV. Given the consistent recommendations for improved surveillance of violence, our findings underscore the utility of using commonly available coded medical data to evaluate services for at-risk groups across the life course. However, estimates varied depending on indicator and outcome, with substantial heterogeneity.

Compared with all other indicators of CM, FAS showed the lowest PPV. This most likely reflects the poor availability of specific ICD-9 codes for FAS, along with the complexity of the diagnosis. All included studies used the ICD-9 code ‘760.71’ focusing on fetuses being affected by alcohol without further description. Yet, the applied reference standard across studies required additional FAS criteria to be met, including facial anomalies and growth deficiencies. There is also an absence of FAS criteria that are widely recognised by clinicians, which may explain a higher underlying rate of misclassifications. Further studies on the accuracy of specific FAS codes using the ICD-10 are needed and may yield higher validity.

The high PPV of coded high-risk injuries of CM aligns with findings from previous reviews based on clinician diagnoses across paediatric healthcare settings. Compared with Kemp et al’s meta-analyses of 32 studies on fractures, our results showed higher PPVs for rib fractures (88.3% vs 70.9%) and similar estimates for lower limb fractures but lower estimates for skull (22.1% vs 30.1%) and upper limb fractures (38.5% vs 47.6%). Our PPV for retinal haemorrhages and burn injuries were also consistent with previous reviews of children across settings. While there are significant methodological differences between this study and previous work (eg, overall age criteria), the consistency between findings suggests that some coded injury patterns could be considered as a broader measure of CM in EHRs to aid identification of high-risk groups.

Coded injury patterns of IPV, such as orbital fractures, provided relatively low PPVs. This is not surprising as the PPV is related to the underlying prevalence of IPV. For example, most studies on IPV were conducted in large populations (eg, EDs) and investigated injury patterns applicable to a wide range of causes. Women are also known to under-report IPV because of safety reasons and stigma. The under-reporting is consistent with the findings that more than 1 in 4 pooled female assault coded records were missing perpetrator information in the underlying medical charts. Still, broader assault-related presentations and upper body contusions showed higher PPVs of IPV (>25%), and their utility in combination with other risk factors might yield comparable predictive accuracy in further studies.

Limitations
This review has important limitations. First, NAS and FAS are not always recognised as forms of prenatal neglect or CM, and the categorisation should be considered with caution to prevent stigma and barriers to help-seeking. Women who misuse substances may be unaware of their pregnancy, and opioids can be prescribed during pregnancy by clinicians for pain or opioid addiction treatment, increasing the risk of NAS.

Second, PPVs were analysed without accounting for study prevalences and within-study correlations of family violence.

50
As a result, pooled estimates from studies with higher underlying prevalence estimates will generally lead to higher PPVs. However, we aimed to minimise variation of underlying prevalences by only including studies from general hospitals or paediatric settings.

Third, we were unable to obtain adequate data to reliably pool estimates on specificity and sensitivity. However, these absolute accuracy measures were not the focus of this review. The high volume of eligible patients presenting to healthcare combined with the rare occurrence and under-reporting of all outcomes limits the feasibility to apply a reference standard to non-coded cases to ascertain false negatives. Reliable measures on sensitivity are therefore unlikely to be obtained.

Fourth, the between-study heterogeneity of the pooled PPVs was substantial. As in most large meta-analyses, several factors not examined in the subgroup analyses may have influenced our results. Many estimates were ascertained from lower quality reference standards, prone to circularity bias (ie, using suspicion of CM as a reference standard in the absence of other explanations). Still, the consistency of studied ICD codes, combined with the larger samples and the high PPV across indicators, suggests that these estimates are valid and merit further study.

Finally, identifying CM and IPV in practice is complex. Obtaining information is a difficult task, as patients often under-report their experiences or symptoms, and high-risk groups such as infants cannot communicate. Similarly, some symptoms addressed by the reference standard (eg, linkage to social service assessments) may not have been conveyed to the clinician. EHRs may thus lead to potentially missed diagnoses or misclassifications. Estimates of indicators, therefore, reflect the ‘best data’ available and should be viewed in terms of routinely recorded indicators to help inform decisions about the likelihood of abuse, rather than definitive diagnoses.

**Implications**

More than a billion children and women aged 0–45 years globally reported being victims of abuse in 2014. Yet, in the UK, studies show that only about 1 in 3 violence-related ED visits for children and adults appear in police records, and self-report studies reveal that 1 in 5 affected women feel reluctant to report abuse to healthcare. In response to WHO’s priorities on addressing gaps in violence prevention, our findings highlight the potential to improve targeted care using routine EHRs to identify, prevent and support high-risk groups of family violence. On a service level, coded indicators of co-occurring family violence have the potential to be incorporated into computerised clinical decision support systems or risk prediction models to flag potential at-risk individuals. In the UK, linkage of family members EHRs could also allow for a ‘Think-Family’ approach, where indicators have the potential to identify vulnerable children through mothers or vice versa. Despite these potential implications, it remains unknown whether the benefits of using automated EHR systems to identify at-risk individuals outweigh potential harms including stigma, legal consequences, trust and reduced help-seeking.

**Acknowledgements**  First, we are very grateful to Dr Erin Shriver (University of Iowa Hospitals and Clinics), Professor John Leventhal (Yale School of Medicine), Dr Jennifer N. Lind (Centers for Disease Control and Prevention), Peggy Brozovic ’ (Vermont Department of Health), Luigi Garcia Saavedra (New Mexico Department of Health), Dr Jane Forro (Illinois Department of Public Health), Dr Ulf Höglund (Uppsala University), Björn Tingberg (Linköping University) and Dr Melissa O’Donnell (The University of Western Australia) for responding to additional data requests to complete our analyses. Second, this research benefits from and contributes to the National Institute for Health Research (NIHR) Children and Families Policy Research Unit but was not commissioned by the NIHR Policy Research Programme. The study also benefited from funding for infrastructure through the NIHR GOSH Biomedical Research Centre and Health Data Research UK. Finally, we express our deep appreciation to the children and families worldwide for their contribution and participation in interpersonal violence research.

**Contributors**  Concept: SS and RG. Design: SS. Drafting of the manuscript: SS, RA, RG and LL. Literature search and screening: SS, RA and MS. Acquisition, analysis or interpretation of data: all authors contributed equally. Statistical analysis: SS and LL. Critical revision of the manuscript for important intellectual content: all authors contributed equally. Study supervision: RG and LL.

**Funding**  The corresponding author had full access to all of the data and had final responsibility to submit for publication.

**Competing interests**  None declared.

**Patient consent for publication**  Not required.

**Provenance and peer review**  Not commissioned; externally peer reviewed.

**Data availability statement**  All data relevant to the study are included in the article or uploaded as supplementary information.

**Open access**  This is an open access article distributed in accordance with the Creative Commons Attribution Non-Commercial (CC BY-NC 4.0) license, which permits others to distribute, remix, adapt, build upon this work non-commercially, and license their derivative works on different terms, provided the original work is properly cited, appropriate credit is given, any changes made indicated, and the use is non-commercial. See: http://creativecommons.org/licenses/by-nc/4.0/.
16 Björkström C, Kosidou K, Björkenstam E. Childhood adversity and risk of suicide: cohort study of 548,721 adolescents and young adults in Sweden. BMJ 2017;357:j3334.

17 Chandran JS, Thomas T, Gokhale KM, et al. The burden of mental illness associated with childhood maltreatment in the UK, using the health improvement network database: a population-based retrospective cohort study. Lancet Psychiatry 2019;6:292–34.

18 Drinkwater J, Stanley N, Szlasy E, et al. Juggling confidentiality and safety: a qualitative study of how general practice clinicians document domestic violence in families with children. Br J Gen Pract 2017;67:e437–44.

19 Herman-Giddens M et al. Underestimation of child abuse mortality in the United States. JAMA 1999;282:463–7.

20 Burns EM, Rigby E, Mamadonna R, et al. Systematic review of discharging coding accuracy. J Public Health 2012;34:138–48.

21 McKenzie K, Scott DA, Waller GS, et al. Reliability of routinely collected hospital data for child maltreatment surveillance. BMC Public Health 2011;11:88.

22 McNees MDF, Moher D, Thoms B, et al. Preferred reporting items for a systematic review and meta-analysis of diagnostic test accuracy studies: the PRISMA-DTA statement. JAMA 2018;320:388–96.

23 Stroup DF, Berlin JA, Morton SC, et al. Meta-Analysis of observational studies in epidemiology: a proposal for reporting. JAMA 2000;283:285–9.

24 Syed S, Ashrick R, Gonzalez-Izquierdo A, et al. Diagnostic accuracy of interpersonal violence and maltreatment in electronic health records: a systematic review and meta-analysis (CRD42019139300) 2019.

25 Covidence systematic review software [program], Melbourne, Australia 2017.

26 Häyrimm K, Saranto K, Nykänen P. Definition NIP: Definition, structure, content, use and impacts of electronic health records: a review of the research literature. Int J Med Inform 2008;77:291–304.

27 Garcia-Moreno C, Janssens H, Ellsberg M, et al. WHO multi-country study on women’s health and domestic violence against women. Geneva: World Health Organization, 2005: 204. 1–18.

28 Whiting P et al. QUADAS-2: a revised tool for the quality assessment of diagnostic accuracy studies. Ann Intern Med 2011;155:529–36.

29 Kemp AM, Dunstan F, Harrison S, et al. Patterns of skeletal fractures in child abuse: systematic review. BMJ 2008;337:a1518.

30 Carlin JB. Meta-Analysis for 2 X 2 tables: a Bayesian approach. Stat Med 1992;11:141–58.

31 Stijnen T, Hamza TH, Øzdemir P. Random effects meta-analysis event outcome in the framework of the generalized linear mixed model with applications in sparse data. Stat Med 2011;30:3906–67.

32 Borenstein M, Higgins JPT, editors. Basics of meta-analysis: I is not an absolute measure of heterogeneity. Res Synth Methods 2017;8:5–18.

33 Sterne JA, Egger M, Moher D. Addressing reporting biases. Cochrane handbook for systematic reviews of interventions: Cochrane book series, 2008: 297–333.

34 Team R Core. R: a language and environment for statistical computing. 2013.

35 Schwarzer G. Meta: an R package for meta-analysis. R News 2007;7:40–5.

36 Alcala ME, Berklehammer SK, Wolden GE, et al. Effects of partially hydrolyzed formula on severity and outcomes of neonatal abstinence syndrome. Am J Perinatol 2019.

37 Chisamore B, Alsaleem M, Berk Schvarzer G, et al. Patterns of skeletal fractures in children younger than three years of age. Eur J Pediatr 2013;172:749–53.

38 Borenstein M, Hedges LV, editors. Basics of meta-analysis: II is not an absolute measure of heterogeneity. Res Synth Methods 2017;8:5–18.

39 Allen K, Riley M, Goldfeld S, et al. Estimating the prevalence of fetal alcohol syndrome in Victoria using routinely collected administrative data. Aust N Z J Public Health 2007;31:62–6.

40 Egeland GM, Perham-Hester KA, Gessner BD, et al. Fetal alcohol syndrome in Alaska, 1977 through 1992: an administrative prevalence derived from multiple data sources. J Public Health 1998;88:781–6.

41 Fox DJ, Druschel CM. Estimating prevalence of fetal alcohol syndrome (FAS): effectiveness of a perinatal birth defects registry system. Birth Defects Res C Mol Teratol 2013;97:604–8.

42 Harris KR, Bucko I. Prevalence of fetal alcohol syndrome in the top end of the Northern Territory. J Paediatr Child Health 2003;39:528–33.

43 Miller L, Tolliver R, Druschel C, et al. Fetal alcohol syndrome–Alaska, Arizona, Colorado, and New York, 1995–1997. MMWR Morb Mortal Wkly Rep 2002;51:433.

44 Miller LA, Shahi T, Stanton C, et al. Surveillance for fetal alcohol syndrome in Colorado. Public Health Rep 1995;110:690.

45 Welty T, Canfield L, Selva K, et al. Use of international classification of diseases coding to identify fetal alcohol syndrome--Indiana Health Service facilities, 1981–1992. MMWR Morb Mortal Wkly Rep 1995;44:253.

46 Anderst JD. Assessment of factors resulting in abuse evaluations in young children with minor head trauma. Child Abuse Negl 2008;32:405–13.

47 Baldwin K, Pandya NK, Wolfgruber H, et al. Femur fractures in the pediatric population: abuse or accidental trauma? JAMA J Public Health Epidemiology: an absolute measure of heterogeneity. J Epidemiol Community Health 2000;54:296–302.

48 Ettaro L, Berger RP, Sonter J. Abusive head trauma in young children: characteristics and medical charges in a hospitalized population. Child Abuse Negl 2004;28:1099–111.

49 Evansovich M, Klein R, Muakkassa F, et al. The economic effect of child abuse in the burn unit. Burns 1998;24:642–5.

50 Garza HH. Diagnostic coding of physical abuse among patients evaluated by a multidisciplinary child protection team in a pediatric level I trauma center 2012.

51 Gessner BD, Moore M, Hamilton B, et al. The incidence of infant physical abuse in Alaska. Child Abuse Negl 2004;28:92–23.

52 Gonzalez-Izquierdo A, Ward A, Smith P, et al. Notifications for child safeguarding from an acute hospital in response to presentations to healthcare by parents. Child Healthcare 1995;44:253.

53 Limb fractures and nonaccidental injury in children less than 24 months of age. Pediatr 2015;134:44–50.

54 Högborg U, Sernnerstam R, Wester K, et al. Medical diagnoses among infants at entry in out-of-home care: a Swedish population-register study. Health Sci Rep 2019;2:e132.

55 Hoft J, Ronda J, Schaeffer P, et al. Identification of physical abuse cases in hospitalized children: accuracy of international classification of diseases codes. J Pediatr 2013;162:80–5.

56 Hoft AM, Asnes AG, Livingston N, et al. The accuracy of ICD codes: identifying physical abuse in 4 children’s hospitals. Acad Pediatr 2015;15:444–50.

57 Hui C, Loughin E, Goldstein S, et al. Femoral fractures in children younger than three years: the role of nonaccidental injury. J Pediatr Orthop 2008;28:297–302.

58 Karatekin C, Almy B, Mason SM, et al. Health-care utilization patterns of maltreated youth. J Pediatr Psychol 2018;43:654–65.

59 Keenan HT, Marshall SW, Nozera MA, et al. Increased incidence of inflicted traumatic brain injury in children after a natural disaster. Am J Prev Med 2003;24:189–93.

60 Kim PT, McCaig L, Dandona A, et al. Consistent screening of admitted infants with head injuries reveals high rate of nonaccidental trauma. J Pediatr Surg 2017;52:1827–30.

61 Krawiec C, Gerard S, Iriana S, et al. What we can learn from failure: an EHR-based child protection alert system. Child maltreatment 2019;1077595919848845.

62 Lane WG, Rubins DM, Montenett R, et al. Racial differences in the evaluation of pediatric fractures for physical abuse. JAMA 2002;288:1603–9.

63 Laskey AL, Stump TE, Hicks RA, et al. Yield of Skeletal Surveys in Children ≤18 Months of Age Presenting with Isolated Skull Fractures. J Pediatr 2013;162:86–9.

64 Lavin LN, Penrod CH, Estrada CA, et al. Fractures in the pediatric emergency department: are we considering abuse? Clin Pediatr 2018;57:1161–7.

65 Lyons TW, Stack AM, Monetaux MC, et al. A QI initiative to reduce hospitalization for children with isolated skull fractures. Pediatrics 2016;137:e20153370.

66 Myhre MC, Greggaard JB, Dyb GA, et al. Traumatic head injury in infants and toddlers. Acta Paediatr 2007;96:1159–63.

67 O’Donnell M, Nassar N, Jacoby P, et al. Western Australian emergency department presentations related to child maltreatment and intentional injury: population level study utilising linked health and child protection data. J Paediatr Child Health 2012;48:57–65.

68 O’Donnell M, Nassar N, Leonard H, et al. Rates and types of hospitalisations for children who have subsequent contact with the child protection system: a population based case-control study. J Epidemiol Community Health 2010;64:784–8.
Intimate partner violence and risk of hospitalization among women. *Arch Dis Child* 2011;96:1042–8.

Ranganath R, Srinivasan RJ, Steinberg RL, et al. Prevalence of retinal hemorrhages in infants presenting with isolated long bone fractures and evaluation for abuse. *J Emerg Med* 2016;51:365–9.

Raghavan R, Brown DS, Allaire BT, et al. Challenges in using Medicaid claims to ascertain child maltreatment. *Child Maltreat* 2015;20:83–91.

Rangel EL, Cook BS, Bennett BL, et al. Eliminating disparity in evaluation for abuse in infants with head injury: use of a screening guideline. *Pediatr Surg* 2009;44:1229–35.

Ricci L, Giantris A, Meriampil P, et al. Abusive head trauma in Maine infants: medical, child protective, and law enforcement analysis. *Child Abuse Negl* 2004;28:271–83.

Risuen SR, Suskaer SJ, Dematt EJ, et al. Functional outcomes in children with abusive head trauma receiving inpatient rehabilitation compared with children with abusive head trauma managed on an inpatient hospital service. *Arch Pediatr Adolesc Med* 2010;164:613–7.

Rynzar E, Rosado N, Flaherty EG. Understanding forensic young children: abuse or not abuse? *Child Abuse Negl* 2015;47:132–9.

Schnitzer PG, Slugger P, Van Tuinen M. Child maltreatment in Missouri: combining data for public health surveillance. *Am J Prev Med* 2004;27:379–84.

Schnitzer PG, Slugger PL, Kruse RL, et al. Identification of ICD codes suggestive of child maltreatment. *Child Abuse Negl* 2011;35:3–17.

Sharky MS, Buesser KE, Gaither JR, et al. Abusive fracture incidence over three decades at a level 1 pediatric trauma center. *Child Abuse Negl* 2018;76:364–71.

Somji Z, Plint A, McGahren C, et al. Diagnostic coding of abuse related fractures at two children’s emergency departments. *Child Abuse Negl* 2011;35:905–14.

Starr RL, Siegel RM, Shapiro RA. Humeral fractures without obvious etiologies in children less than 3 years of age: when is it abuse? *Pediatrics* 1995;96:667–71.

Thyen U, Leventhal JM, Yazgideri SR, et al. Concerns about child maltreatment in hospitalized children. *Child Abuse Negl* 1997;21:187–98.

Tingberg B. Child abuse-clinical investigation, management and nursing approach 2010.

Tzioumi D, Orates RK. Subdural hematomas in children under 2 years: accidental or inflicted? A 10-year experience. *Child Abuse Negl* 1998;22:1105–12.

Wood JN, Christian CW, Adams CM, et al. Skeletal surveys in infants with isolated skull fractures. *Pediatrics* 2009;123:e247–52.

M-FW, T-H L, Lin C-J, et al. Risk factors and physical signs of child abuse in hospitalized children in Taiwan. *Children and Youth Services Review* 2015;58:137–41.

Durand MB, McLaughlin CM, Imagawa KK, et al. Identifying targets to improve coding of child physical abuse at a pediatric trauma center. *Journal of Trauma Nursing* 2019;26:239–42.

Wood JN, Pecker LH, Russo ME, et al. Evaluation and referral for child maltreatment in pediatric poisoning victims. *Child Abuse Negl* 2012;36:362–9.

Aroseroa OA, Fritsch TA, Hueso Y, et al. Maxillofacial injuries and violence against women. *Arch Facial Plast Surg* 2009;11:48–52.

Biroscak BJ, Smith PK, Rozzwecki H, et al. Intimate partner violence against women: findings from one state’s ED surveillance system. *J Emerg Nurs* 2006;32:12–16.

Bonomi AEF et al. Medical and psychosocial diagnoses in women with a history of intimate partner violence. *Arch Intern Med* 2009;169:1692–7.

Brown S, Malose L, Carson E, et al. Intimate partner violence injuries—Okahoma, 2002. *MMWR Morb Mortal Wkly Rep* 2005;54:1041–5.

Clark RJ, Rehe LM, Soto JA, et al. Intimate partner violence: an underappreciated etiology of orbital floor fractures. *Ophthalmic Plast Reconstr Surg* 2014;30:508–11.

Cohen AR, Clark TJ, Renner LM, et al. Intimate partner violence as a mechanism of traumatic ocular injury in women. *Can J Ophthalmol* 2019;54:355–8.

Farislow JL, Norton RN, Spinola CG. Indicators of assault-related injuries among women presenting to the emergency department. *Ann Emerg Med* 1998;32:341–8.

Goldberg SN, McCann C. Three years of age: when is it abuse? Pediatrics 1995;96:667–71.

Kivela S, Leppakoski T, Ruohoniemi J, et al. The documentation and characteristics of hospitalized IPV patients using electronic medical records data: a follow-up descriptive study. *J Fam Violence* 2019;34:611–9.

Lipsky S, Caetano R, Roy-Byrne P. Racial and ethnic disparities in police-reported intimate partner violence and risk of hospitalization among women. *Women’s Health Issues* 2009;19:109–18.

Muellerlen RL, Lenaghan PA, Pakesier RA. Battered women: injury locations and types. *Ann Emerg Med* 1996;28:486–92.