An Extension of Generalized Triphasic Logistic Human Growth Model

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Abstract
The purpose of the present study was to establish an extension of generalized triphasic logistic human growth model. This model was applied through the higher dimensional growth process. Bayesian method of estimation was applied to estimate the parameters of the model. Principal component regression method was applied when there was a problem of multicollinearity in the model. The biological characteristics of the human growth process were extracted from the velocity and acceleration curve using gradient vector and Hessian matrix.

Keywords: Higher dimensional growth model; Human biological characteristics; Logistic model; Principal component regression; Bayesian approach; Velocity; Acceleration; Gradient vector; Hessian matrix

Introduction
Growth modeling is very important in recent time for the purpose of the study of growth pattern and biological characteristics. Growth in stature in any living organism depends on their: age, body weight, chest circumference, sitting height, genetic factors, maternal illnesses during pregnancy, socio-economic disadvantages during and after pregnancy, social/emotional problems during childhood, poor nutrition and environmental or emotional deprivation and so on. Most of the growth models such as Gompertz and the logistic growth model [1,2], Jenss model [3], Count model [4,5], Double logistic model [6], PB models [7,8], ICP model [9], Reed models [10], SSC model [8,11], JPPS model [12], JPA-1 and JPA-2 model [13], Modified ICP model [14], BTT model [15] have considered stature or weight as a function of age. Comparative study among different fitted growth models was done by many authors [8,16,17]. The JPA-2 [13] model fits better than all other asymptotic models till 1991. While, BTT [18] model is found to be better than JPA-2 model. Like age, it is important to incorporate other possible predictors in the model as they have significant influence on stature, however, the influence of age on stature is higher than others predictors.

Thus, for the better explanation of the final stature, higher dimensional growth model should be applicable. Rahman et al. [19] showed that the BTT model fits better for those who have early, middle and adolescent growth spurts.

Thus, the purpose of the present study is to develop a higher dimensional growth model through the extension of BTT model.

History of the BTT Model
Robertson [20] proposed that the human growth of organism in general occurs in a number of additive, more-or-less independent phases during the course of development. Generally, the timing and intensity of each phase is assumed genetically programmed in the individual, but they may vary in expression according to environmental conditions. According to Robertson, the sum of the three logistic components can be describing human growth in stature but he cannot test the model because of no suitable data. Bock and Thissen (1976) [21], first applied the concept to individual growth using the case from the Berkeley and Fels growth studies. Their analysis showed that the goodness of fit of the triphasic logistic model was good over the range from one year to maturity. In the Bock-Thissen model, the phases represent early-childhood, middle-childhood, and adolescent growth. A further enhancement of the Bock-Thissen model was suggested by du Toit [22]. du Toit [22] suggested the additional of the “shape” constants of the positive exponentials to the denominators of the logistic functions to control the model in the region of change-over from early to middle childhood, and to provide some asymmetry of the adolescent component. We found values for these constants that tend to improve the fit of the model, by repeated trials with the Berkeley and Fels data. We refer to this triphasic generalized logistic model as the BTT model.

Mathematical Explanation of the BTT Model
Bock-Thissen-du Toit (BTT) model can be defined as sum of the three generalized logistic terms. The form of the logistic term is:

\[ \frac{a}{1 + e^{-(bt+c)}} \]

where, \( t \) is the time (age) variable; \( a, b, c \) and \( d \) are the amount of growth, slope, intercept and fixed shape constant contributed by the term, respectively. The quantity \( z=(bt+c) \) in the exponential function is the “logit”.

Bock et al. [15] described a model known as the triphasic generalized logistic model by summing up three phases of growth; early, middle and adolescent. This triphasic generalized logistic model can be written as follows:

\[ y = \frac{a_1}{1 + e^{-(bt+c)}} + \frac{a_2}{1 + e^{-(bt+c)}} + \frac{a_3}{1 + e^{-(bt+c)}} \]

where, the set of parameters \((a_1, b_1, c_1)\), \((a_2, b_2, c_2)\) and \((a_3, b_3, c_3)\) refer to the parameters of early, middle and adolescent phases of growth, respectively.

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Proposed Model

Total body growth should be dependent on many factors e.g., age, body weight, chest circumference, sitting height, genetic factors, and maternal illnesses during pregnancy and so on. To the best of our knowledge, all growth models are proposed on age factor only (Gompertz and the logistic growth model, Jenss model, Count model, ICP model, JPPS model, JPA-1 and JPA-2 model, Modified ICP model and BTT model). We have taken here the opportunity to incorporate all other factors in the BTT growth model. Why we have used BTT model? Because very recently, it is found that BTT growth model perform well than all other parametric growth models [19].

Because no problem of multicollinearity.

We can write the equation (1) as follows:

\[ y = \frac{a_1}{1 + e^{-\left(a_1 x_1 + a_2 x_2 + \ldots + a_p x_p + c_1\right)}} + \frac{a_2}{1 + e^{-\left(a_1 x_1 + a_2 x_2 + \ldots + a_p x_p + c_2\right)}} + \frac{a_3}{1 + e^{-\left(a_1 x_1 + a_2 x_2 + \ldots + a_p x_p + c_3\right)}} \ (1) \]

where, \( y, x_1, x_2, \ldots, x_p \) are the different measurement of body which may be stature, age, weight, chest circumference, sitting height and so on, respectively. The set of parameters \( (a_1, a_2, a_3, \ldots, a_p, c_1, c_2, c_3) \), \( (a_1, a_2, a_3, \ldots, a_p, c_1) \) and \( (a_1, a_2, a_3, \ldots, a_p, c_1) \) refer to the parameters of early, middle and adolescent phases of growth in Euclidian space, respectively. Bock and Thissen (1980) [23] imposed a linear restriction on the parameters of the first and second term to remove the over parameterization problem, but du Toit [22] later found that setting \( c_i = 0 \) serves equally well.

To estimate proposed model, there are two cases, such as:

The regressors are uncorrelated, and

The regressors are correlated.

Case (1): The regressors are uncorrelated

When the regressors are uncorrelated, we can estimate parameters of equation (1) directly by Bayesian approach described in Estimation section because there is no problem of multicollinearity.

Case (2): The regressors are correlated

Principal components regression [24] is applied in case of correlated regressors. We can write the equation (1) as follows:

\[ y = \frac{a_1}{1 + e^{-\left(x_1 a_1 + x_2 a_2 + \ldots + x_p a_p + c_1\right)}} + \frac{a_2}{1 + e^{-\left(x_1 a_1 + x_2 a_2 + \ldots + x_p a_p + c_2\right)}} + \frac{a_3}{1 + e^{-\left(x_1 a_1 + x_2 a_2 + \ldots + x_p a_p + c_3\right)}} \]

where,

\[ X = (x_1, x_2 \ldots x_p), \]

\( \beta_i = (c_1, a_{11}, a_{12} \ldots a_{1p})' ; \)

\( \beta_2 = (c_2, a_{21}, a_{22} \ldots a_{2p})' ; \)

and

\( \beta_3 = (c_3, a_{31}, a_{32} \ldots a_{3p})' ; \)

For \( n \) data,

\[ X = \begin{bmatrix} 1 & X_{11} & X_{12} & \ldots & X_{1n} \\ 1 & X_{21} & X_{22} & \ldots & X_{2n} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & X_{p1} & X_{p2} & \ldots & X_{pn} \end{bmatrix}_{nx(p+1)} \]

Let, \( \Lambda = \begin{bmatrix} \lambda_1 & \lambda_2 & \lambda_3 & \ldots & \lambda_p \end{bmatrix} \) is a \( (p+1) \times (p+1) \) diagonal matrix of the eigenvalues of \( XX' \) and \( T = T'X \) is a \( (p+1) \times (p+1) \) orthogonal matrix whose columns are the eigen vectors associated with \( \lambda, \lambda, \lambda, \ldots, \lambda \).

We can define a new set of orthogonal regressors, such as \( Z = (Z_1, Z_2, \ldots, Z_p) \) which is the column of \( Z \) are referred to as principal components.

Therefore,

\[ y = \frac{a_1}{1 + e^{-\left(Z_1 a_1 + Z_2 a_2 + \ldots + Z_p a_p + c_1\right)}} + \frac{a_2}{1 + e^{-\left(Z_1 a_1 + Z_2 a_2 + \ldots + Z_p a_p + c_2\right)}} + \frac{a_3}{1 + e^{-\left(Z_1 a_1 + Z_2 a_2 + \ldots + Z_p a_p + c_3\right)}} \]

The principal components regression approach used less than full set of principal components to combat multicollinearity in the model. In principal components estimators, we assume that the regressors are arranged in order of decreasing eigenvalues:

\[ \lambda_0 \geq \lambda_1 \geq \lambda_2 \geq \ldots \geq \lambda_p > 0 \]

Let us suppose that the last \( s \) of these eigenvalues approximate equal to zero. In principal components regression the principal components corresponding to near-zero eigenvalues are removed from the analysis and Bayes estimate defined in Estimation section is applied to the remaining components. That is,

\[ y_{pc} = B y_i, i = 1, 2, 3 \]

where, \( b_1 = b_2 = \ldots = b_{p-s} = 1 \) and \( b_{p-s+1} = b_{p-s+2} = \ldots = b_p = 0 \)

and the fitted model can be written as:
\[ \hat{y} = \frac{a_1}{1 + e^{(-X \beta_1)}} + \frac{a_2}{1 + e^{(-X \beta_2)}} + \frac{a_3}{1 + e^{(-X \beta_3)}} \]

Replacing \( Z \) by the linear combination of \( X \), we get

\[ \hat{y} = \frac{a_1}{1 + e^{(-X \beta_1)}} + \frac{a_1}{1 + e^{(-X \beta_1)}} + \frac{a_1}{1 + e^{(-X \beta_1)}} \]

Thus, the principal component estimator can be written as follows:

\[ \hat{\beta}_p = T \hat{\beta}_p \quad \text{for } i = 1, 2, 3 \quad \text{[since } \hat{y}_i = T' \beta_i \text{ for } i = 1, 2, 3 \text{ and } TT = I_{p+1}] \]

**Estimation**

Before estimating the biological parameters of proposed model, we fixed up the shape parameters. The shape parameters can be estimated by trial and error methods. One way is to fix the shape parameters such that the value of shape parameter of third phase must be greater than any other parameter. Generally, the value of shape parameter of third phase must be greater than the other two phases. Generally, the value of shape parameters of the first two phases is equal. Now let us consider the proposed nonlinear growth model of the form:

\[ y = f(X, \Theta) = \sum_{i=1}^{3} \frac{a_i}{1 + e^{(-X \beta_i + \gamma_i)}} \]

(2)

where, \( y \) is the dependent variable, \( X = (x_1, x_2, x_3, x_4) \) is the vector of independent variables, \( \Theta = (a_1, a_2, a_3, \beta_1, \beta_2, \gamma_1, \gamma_2, \gamma_3) \) and \( \gamma_i \) is the parameters vector of the model. When observations of \( y \) and \( X \) are collected for observation \( i \), the equation (2) becomes:

\[ y_i = f(X_i, \Theta) + \epsilon_i \quad i = 1, 2, 3 \ldots n \]

(3)

where, the model error \( \epsilon_i \sim \mathcal{N}(0, \sigma^2) \) distribution. Since, in this study the number of distinct observations is greater than or equal to number of parameters, and hence the estimation by the conventional least squares method for complex growth models is less than the ideal. Even when the number of observations is sufficient for least squares, the parameters may not all be identifiable if the observations are poorly positioned. A much better method for fitting growth models is Bayes model estimation which chooses among a specified population of growth curves.

The random vector of parameters \( \Theta \) is assumed to follow \( \mathcal{N}(\mu, \sigma^2) \) distribution in the population. Let us consider a squared error loss function:

\[ l(\Theta, \hat{\Theta}) = (\Theta - \hat{\Theta})^2 \]

Since, \( \epsilon \sim \mathcal{N}(0, \sigma^2) \), then

\[ E(y | X, \Theta) = f(X; \Theta) + E(\epsilon) = f(X; \Theta) + 0 \]

\[ \therefore E(y | X, \Theta) = f(X; \Theta), \quad \text{and} \quad V(y | X, \Theta) = \sigma^2 \]

So that, \( y | X, \Theta, \sigma^2 \sim \mathcal{N}(f(X; \Theta), \sigma^2) \)

We consider a sample \( (y_1, y_2, y_3, \ldots, y_n) \) of size \( n \) from the density \( f(y | X, \Theta, \sigma^2) \). Then the likelihood function is defined as follows:

\[ L(y | X, \Theta, \sigma^2) = \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(y_i - f(X_i; \Theta))^2} \]

\[ \therefore L(y | X, \Theta, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(y_i - f(X_i; \Theta))^2} \]

Posterior Bayes estimator of \( (\Theta, \sigma^2) \) with respect to prior \( g(\Theta, \sigma^2) \) can be written as:

\[ \pi(\Theta, \sigma^2 | y) = \frac{L(y | X, \Theta, \sigma^2)g(\Theta, \sigma^2)}{\int_{\Theta} L(y | X, \Theta, \sigma^2)g(\Theta, \sigma^2) d\Theta d\sigma^2} \]

The denominator in the above posterior distribution is constant. Thus, the posterior distribution can be represented by the form:

\[ \pi(\Theta, \sigma^2 | y) \propto L(y | X, \Theta, \sigma^2)g(\Theta, \sigma^2) \]

(4)

The most important part in the Bayesian regression analysis is to determine the prior distribution. However, it is very difficult to infer the probability distribution of the regression coefficient in the separated basins. Thus, a short of uniform prior distribution is selected to compute the posterior distribution in the study. Sorensen and Gianola [25] suggested a sort of uniform distribution using variance:

\[ g(\Theta, \sigma^2) \propto \frac{1}{\sigma^2} \]

(5)

From (4) and (5), we can write,

\[ \pi(\Theta, \sigma^2 | y) \propto \frac{1}{\sqrt{\sigma^2}} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^{n} (y_i - f(X_i; \Theta))^2} \]

\[ \therefore \pi(\Theta, \sigma^2 | y) \propto (\sigma^2)^{-\frac{n}{2}-1} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^{n} (y_i - f(X_i; \Theta))^2} \]

We know that, for a squared error loss function, the mean of the posterior density is the Bayes estimator of \( \Theta \).

Now, the Bayes estimator of under squared error loss function is

\[ \hat{\Theta} = E(\pi(\Theta)) = \int_{\Theta} \Theta \pi(\Theta, \sigma^2) d\Theta \]

\[ \therefore \hat{\Theta} = \int_{\Theta} (\Theta | \sigma^2)^{\frac{n}{2}+1} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^{n} (y_i - f(X_i; \Theta))^2} d\Theta \]

The solution of above equation (6) can be determined numerically. Fisher-scoring (Newton-Gauss) method is extremely fast, and nearly as robust as MEAP (Minimum Expected a Posteriori) estimation.

**Velocity**

Velocity is a term used for a rate of change. That is, velocity is defined as the ratio of the directed displacement \( \Delta r \) (say) to the required time \( \Delta t \) (say). That is,
Velocity = \( \frac{\Delta r}{\Delta t} \)

The direction of \( \Delta r \) is the same as the direction of the displacement.

\[ y = f(X, \Theta) = \sum_{i=1}^{\infty} \frac{a_i}{\left[1 + e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)}\right]^2} \]

Velocity for the variable \( X_1 \)

\[ \frac{\partial y}{\partial x_1} = \sum_{i=1}^{\infty} \frac{a_i d_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)}}{\left[1 + e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)}\right]^2} \]

Velocity for the variable \( X_2 \)

\[ \frac{\partial y}{\partial x_2} = \sum_{i=1}^{\infty} \frac{a_i d_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)}}{\left[1 + e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)}\right]^2} \]

Velocity for the variable \( X_3 \)

\[ \frac{\partial y}{\partial x_3} = \sum_{i=1}^{\infty} \frac{a_i d_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)}}{\left[1 + e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)}\right]^2} \]

Similarly, Velocity for the variable \( X_p \)

\[ \frac{\partial y}{\partial x_p} = \sum_{i=1}^{\infty} \frac{a_i d_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)}}{\left[1 + e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)}\right]^2} \]

In general

\[ \frac{\partial y}{\partial x_j} = \sum_{i=1}^{\infty} \frac{a_i d_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)}}{\left[1 + e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)}\right]^2} \]

\( j = 1, 2, 3, \ldots p \)

**Acceleration**

Acceleration of a function \( y = f(X, \Theta) \) is defined as the second derivative of \( y \) with respect to \( x \). That is,

\[ \text{Acceleration} = \frac{\partial^2 y}{\partial x^2} \]

Acceleration for the variable \( X_1 \)

\[ \frac{\partial^2 y}{\partial x_1^2} = \sum_{i=1}^{\infty} \left[ a_i d_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)} + a_i d_i a_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)} \right] \]

Acceleration for the variable \( X_2 \)

\[ \frac{\partial^2 y}{\partial x_2^2} = \sum_{i=1}^{\infty} \left[ a_i d_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)} + a_i d_i a_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)} \right] \]

Acceleration for the variable \( X_3 \)

\[ \frac{\partial^2 y}{\partial x_3^2} = \sum_{i=1}^{\infty} \left[ a_i d_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)} + a_i d_i a_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)} \right] \]

\( \vdots \)

Acceleration for the variable \( X_p \)

\[ \frac{\partial^2 y}{\partial x_p^2} = \sum_{i=1}^{\infty} \left[ a_i d_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)} + a_i d_i a_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)} \right] \]

In general

\[ \frac{\partial^2 y}{\partial x_j^2} = \sum_{i=1}^{\infty} \left[ a_i d_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)} + a_i d_i a_i a_i e^{-(a_1 x_1 + a_2 x_2 + a_3 x_3 + \ldots + a_p x_p + c_i)} \right] \]

\( j = 1, 2, 3, \ldots p \)

To compute the Hessian matrix, we need various derivative of \( y = f(X, \Theta) \) with respect to:

\( x_1, x_2, \ldots, x_j \), and \( x_k \) as follows:

Now,
Partial derivative of y with respect to the variable $x_i$ and $x_j$:

$$
\frac{\partial^2 y}{\partial x_i \partial x_j} = \sum_{k=1}^{3} \left[ \frac{a d_{i}^{2} a_{j} a_{k} e^{-\left(a x_{i} + a x_{j} + a x_{k} + \ldots + a x_{p} + \alpha_{x} \gamma_{y} \right)}}{1 + e^{-\left(a x_{i} + a x_{j} + a x_{k} + \ldots + a x_{p} + \alpha_{x} \gamma_{y} \right)}} \right] i^{2}
$$

Where $a d_{i}^{2}$ is the second partial derivative of $a_{i}$ with respect to $x_{i}$.

Partial derivative of y with respect to the variable $x_i$ and $x_j$:

$$
\frac{\partial^2 y}{\partial x_i \partial x_j} = \sum_{k=1}^{3} \left[ \frac{a d_{i} a_{j} a_{k} e^{-\left(a x_{i} + a x_{j} + a x_{k} + \ldots + a x_{p} + \alpha_{x} \gamma_{y} \right)}}{1 + e^{-\left(a x_{i} + a x_{j} + a x_{k} + \ldots + a x_{p} + \alpha_{x} \gamma_{y} \right)}} \right] i^{2}
$$

Method of Finding Velocity and Acceleration

Gradient of y

The vector gradient [26] of a function of $p$ variables $f(x_1, x_2, x_3,..., x_p)$ is defined as follows:

$$
\nabla f(x) = \left( \frac{\partial f}{\partial x_1}, \frac{\partial f}{\partial x_2}, ..., \frac{\partial f}{\partial x_p} \right)^T
$$

where, $T$ means transpose of the vector.

Hessian matrix of y

In mathematics, the Hessian matrix is the square matrix of second-order partial derivatives of a function. The Hessian matrix describes the local curvature of a function of many variables. The Hessian matrix was developed in the 19th century by the German mathematician Ludwig Otto Hesse and later named after him. Hesse himself had used the term "functional determinants".

Given the real-valued function:

$$
y = f(X, \Theta) = \sum_{i=1}^{3} \frac{a_{i}}{1 + e^{-\left(a x_{i} + a x_{j} + a x_{k} + \ldots + a x_{p} + \alpha_{x} \gamma_{y} \right)}}
$$

If all second partial derivatives exist, then the Hessian matrix can be written as follows:

$$
\nabla^2 f(x) = \left[ \begin{array}{ccc}
\frac{\partial^2 f}{\partial x_1 \partial x_1} & \frac{\partial^2 f}{\partial x_1 \partial x_2} & \cdots & \frac{\partial^2 f}{\partial x_1 \partial x_p} \\
\frac{\partial^2 f}{\partial x_2 \partial x_1} & \frac{\partial^2 f}{\partial x_2 \partial x_2} & \cdots & \frac{\partial^2 f}{\partial x_2 \partial x_p} \\
\vdots & \vdots & \ddots & \vdots \\
\frac{\partial^2 f}{\partial x_p \partial x_1} & \frac{\partial^2 f}{\partial x_p \partial x_2} & \cdots & \frac{\partial^2 f}{\partial x_p \partial x_p}
\end{array} \right]
$$

Let us consider any value of $x_1, x_2, x_3,..., x_p$ is $x_1, x_2, x_3,..., x_p$, respectively. Then $\nabla^2 f(x_1, x_2, x_3,..., x_p) = 0$ and satisfy any one of the following condition:

Case 1: $f(x_1', x_2', x_3',..., x_p')$ have a local minimum value if:

$$
\frac{\partial^2 f}{\partial x_1 \partial x_1} > 0, \frac{\partial^2 f}{\partial x_2 \partial x_2} > 0, ..., \frac{\partial^2 f}{\partial x_p \partial x_p} > 0
$$

Case 2: $f(x_1', x_2', x_3',..., x_p')$ have a local maximum value if:

$$
\frac{\partial^2 f}{\partial x_1 \partial x_1} < 0, (-1)^{i} \frac{\partial^2 f}{\partial x_i \partial x_i} > 0, (-1)^{i} \frac{\partial^2 f}{\partial x_i \partial x_j} > 0
$$
Case 3: There is no local maxima or minima if \( f(x_1, x_2, x_3, \ldots, x_n) \) does not satisfy the two conditions described above, and called saddle point.

**Numerical Example**

Consider the longitudinal data of Japanese girl from 1 year to 18 year. The data contains the variables age, stature and chest circumference. We used the STATISTICA software to estimate the parameters and the MatLab software help us to find the velocity and acceleration with respect to different variables. Let us consider the case-1, when the regressors are uncorrelated. We have used the study variables such as age and stature for BTT, and age, stature and chest circumference for same shape constants. Therefore, the proposed model is more accurate than BTT model because it explains stature efficiently with small error compared with BTT model. The plot of final loss function corresponding to BTT and EBTB (Proposed) models are shown in figure 1. From figure 1, it is clear that within the acceptable region of shape constants, the loss function for EBTB model is always smaller than BTT model.

The default values of the shape constants given in the Auxal software [15] were also applicable for Japanese population [27]. Thus the values of \( d_1, d_2 \) and \( d_3 \) are considered as 0.75, 0.75 and 1.20, respectively. Therefore, the velocity and acceleration for the proposed model with shape constants of \( d_1=0.75, d_2=0.75 \) and \( d_3=1.20 \) are presented in table 1. The estimated parameters of the proposed model is presented last row of table 1.

**Conclusion**

The proposed extension of BTT model is very important because the stature of human depends on several factors. In BTT model, we only know the impact of age on stature but our proposed model study knows the impact of several factors. There are more than one variable in the regressors, so the multicollinearity problem may arise. In this case, we can apply principal component regression method for regressor’s variables and find a few components. The advantage of PCA is that the

### Table 1: Estimated parameters and final loss for BTT and proposed models.

| Parameter Model | \( d_1 \) | \( d_2 \) | \( d_3 \) | \( \hat{a}_{11} \) | \( \hat{a}_{12} \) | \( \hat{a}_2 \) | \( \hat{a}_{21} \) | \( \hat{a}_{22} \) | \( \hat{c}_1 \) | \( \hat{c}_2 \) | \( \hat{c}_3 \) | Final loss |
|-----------------|--------|--------|--------|--------------|--------------|---------|----------|----------|-------|--------|--------|----------|
| BTT             | 0.75   | 0.75   | 0.75   | 0.81         | -9.10        | 140.63  | 0.27     | -0.55    | -2.62 | 2.23   | -18.77 | 5.01     |
| EBTB            | 0.75   | 0.75   | 0.75   | -10.84      | 52.05        | 1.06    | -0.16    | -0.36    | 93.59 | 1.04   | -0.18  | 8.24     | 18.28   | 1.45     | 0.08    | 1.16   |
| BTT             | 0.50   | 0.75   | 0.75   | -32.24      | -0.94        | 9.65    | 118.58   | 0.34     | -1.13 | 78.94  | -0.04  | 0.72     | 4.26    |
| EBTB            | 0.50   | 0.75   | 0.75   | 54.68       | 1.07         | -0.15   | 98.54    | 0.61     | -0.04 | 1.91   | 10.63  | 2.56     | -0.21   | -0.85   | 1.29    |
| BTT             | 0.60   | 0.80   | 0.80   | 40.42       | -3.03        | 126.19  | 90.58    | 0.40     | -1.23 | 33.20  | 0.77   | -8.76    | 4.62    |
| EBTB            | 0.60   | 0.80   | 0.80   | 45.49       | 0.69         | 0.32    | -31.18   | 23.04    | 1.80  | -1.62 | 91.65  | 119.18  | 0.33    | -0.01   | -0.31   | 2.15   |
| BTT             | 0.70   | 0.80   | 0.90   | -32.20      | -0.82        | 8.72    | 106.43   | 0.35     | -1.12 | 82.56  | -0.03  | 1.21     | 4.37    |
| EBTB            | 0.70   | 0.80   | 0.90   | 36.07       | 0.96         | -0.05   | -7.63    | 83.08    | 0.48  | -0.02  | -0.32 | 45.42    | 1.69    | -0.92   | 48.70   | 2.88   |
| BTT             | 0.75   | 0.75   | 1.20   | 90.05       | 0.41         | -1.32   | 34.20    | 0.77     | -8.86 | 39.90  | 1.84   | 62.41    | 4.63    |
| EBTB            | 0.75   | 0.75   | 1.20   | 74.12       | 1.23         | -0.47   | 22.04    | 52.70    | 0.34  | 0.12   | -7.73  | 37.35    | 0.91    | -0.07   | -4.54   | 3.29   |

Note: BTT and EBTB means Bock-Thissen-du Toit and Extension of Bock-Thissen-du Toit model, respectively.
principal components are independent of each other. The independent case and after applying the principal component method for dependent case thereafter all estimate parameters by Bayesian approach. The gradient vector and Hessian matrix used for finding velocity and acceleration. Since the proposed model is especially the extension of BTT model. So, if we use only one regressor (p=1) by age then the proposed model fall in BTT model.

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References

1. Deming J (1957) Application of Gompertz curve to the observed pattern of growth in length of 48 individual boys and girls during the adolescent cycle of growth. Hum Biol 29: 83-122.
2. Merrell M (1931) The relationship of individual growth to average growth. Hum Biol 3: 37-70.
3. Jenss RM, Bayley N (1937) A mathematical method for studying the growth of a child Hum Biol 9: 19-23.
4. Count EW (1942) A quantitative analysis of growth in certain human skull dimensions Hum biol 14: 143-165.
5. Count EW (1943) Growth patterns of human physique: an approach to kinetic anthropology. Hum biol 15: 1-32.
6. Bock RD, Wainer H, Peterson A, Thissen D, Murray J (1973) A parameterization for individual human growth curve. Hum Biol 45: 63-80.
7. Preece MA, Baines MJ (1978) A new family of mathematical models describing the human growth curve. Ann Hum Biol 5: 1-24.
8. Ledford AW, Cole TJ (1998) Mathematical models of growth in stature throughout childhood. Ann Hum Biol 25: 101-115.
9. Karlberg J (1989) A biologically-oriented mathematical model (ICP) for human growth, Acta Paediatr Scand Suppl 350: 70-94.
10. Berkley CS, Reed RB (1987) A model for describing normal and abnormal growth in early childhood. Hum Biol 59: 973-987.
11. Shohoji T, Sasaki H (1987) Individual growth of Japanese. Growth 51: 432-450.
12. Jolicoeur P, Portier J, Pemin MO, Sempe M (1988) A lifetime asymptotic growth curve for human height. Biometrics 44: 995-1003.
13. Jolicoeur P, Portier J, Aridi H (1992) Asymptotic models for the longitudinal growth of human stature. Am J Hum Biol 4: 461-468.
14. Johnson ML (1993) Analysis of serial growth data. Am J Hum Biol 5: 633-640.
15. Bock RD, Stephen HC, du Toit SHC, Thissen D (2003) AUXUL: Auxological analysis of longitudinal measurements of human stature. Scientific software international.
16. Kanefuji K, Shohoji T (1990) On a growth model of human height. Growth, Development and Ageing 54: 155-166.
17. Shohoji T, Kanefuji K, Sumiya T, Qin T (1991) A prediction of individual growth of height according to an empirical Bayesian approach. Ann Inst Statist Math 43: 607-619.
18. Jolicoeur P, Abidi H, Portier J (1991) Human stature: which growth model. Growth Dev Aging 55: 129-132.
19. Rahman JAMS, Ali MA, Ashizawa K, Ohtsuki F (2004) Prediction of adult stature for Japanese population: an improvement of Ali-Ohtsuki equations. Anthropological science 112: 61-66.
20. Robertson TB (1908) On the normal rate of growth of an individual, and its biochemical significance. Archiv für Entwicklungs Mechanik den Organismen 25: 581-614.
21. Bock RD, Thissen D (1976) Fitting multi-component models for growth in stature. Proceeding of the 9th International Biometrics Conference, Biometric Society, Boston.
22. du Toit (1992) Multi-component models for non-consecutive data with ARMA(1,1) residuals. Department of Statistics, University of Pretoria.
23. Bock RD, Thissen D (1980) Statistical problem of fitting individual growth curves. In: Johnston FE, Roche AF, Susanne C (eds.), Human physical growth and maturation: methodologies and factors, New York 265-290.
24. Izenman AJ (2008) Modern multivariate statistical techniques-regression, classification, and manifold learning. (1stedn), Springer.
25. Sorensen D, Gianola D (2002) Likelihood, Bayesian, and MCMC Methods in Quantitative Genetics. Springer-Verlag, New York.
26. Hyvarinen A, Karhunen J, Oja E (2001) Independent component analysis. John Wiley & Sons Inc.
27. All MA, Ohtsuki F (2001) Prediction of adult stature for Japanese population: A stepwise regression approach. Am J Hum Biol 13: 316-322.