Comparison of the prognostic values of three calculation methods for echocardiographic relative wall thickness in acute decompensated heart failure

Satoshi Yamaguchi1,2, Michio Shimabukuro3, Masami Abe1, Tomohiro Arakaki1, Osamu Arasaki1 and Shinichiro Ueda2*

Abstract

Purpose: Left ventricular (LV) wall thickness can be measured at the posterior wall (PW) and the intraventricular septum (IVS) in a parasternal long axis view by transthoracic echocardiography. Thus, there are three methods to calculate relative wall thickness as follows: $RWT_{PW} = 2 \times PWth/LVDd$; $RWT_{IVS+PW} = (IVSth + PWth)/LVDd$; and $RWT_{IVS} = 2 \times IVSth/LVDd$ (IVSth = interventricular septum thickness; LVDd = LV internal dimension at end–diastole; PWth = posterior wall thickness). The aim was to compare the prognostic values of these RWTs in patients with acute decompensated heart failure (ADHF).

Method: This was a single-center, retrospective, observational study at a Japanese community hospital. A total of 389 hospitalized ADHF patients were divided into two groups based on the three median RWT values. The primary outcome was all-cause death. Survival analysis was performed, and Cox proportional hazard models unadjusted and adjusted by Get With The Guideline score were used.

Results: High-RWTPW had poor survival (log-rank, $P = 0.009$) and was a significant risk (unadjusted HR (95%CI), 1.72 (1.14–2.61), $P = 0.01$; adjusted HR, 1.95 (1.28–2.98), $P = 0.02$). High-RWTIVS + PW was not associated with poor survival on survival analysis or the unadjusted Cox model. Only the adjusted Cox model showed that High-RWTIVS + PW was associated with a significant risk of the primary outcome (unadjusted HR (95%CI), 1.45 (0.96–2.17), $P = 0.07$; adjusted HR, 1.53 (1.01–2.32), $P = 0.045$). High-RWTIVS did not have significant prognostic value.

Conclusions: When calculating RWT, RWTPW should be recommended for evaluating the mortality risk in ADHF.

Keywords: Concentric left ventricular structure, Relative wall thickness, Acute decompensated heart failure, Transthoracic echocardiography, Prognosis

Introduction

A concentric left ventricular (LV) structure is the result of remodeling that occurs with LV wall thickening relative to the LV cavity to compensate for pressure overload [1, 2]. A concentric LV structure is a risk factor for cardiovascular events in hypertensive patients [3, 4]. Furthermore, we previously reported that a concentric LV structure evaluated by transthoracic echocardiography (TTE) was associated with poor survival in patients with acute decompensated heart failure (ADHF) [5].

Relative wall thickness (RWT) is an index of LV concentricity. RWT is the ratio of LV wall thickness to the LV internal dimension at end diastole (LVDd) [6]. LV wall thickness, which can be measured in a parasternal long-axis view by TTE, is represented by the Interventricular septum wall thickness (IVSth) and the posterior wall thickness (PWth) [6]. Therefore, there are three methods to calculate the RWT: $RWT_{PW} = 2 \times PWth/LVDd$; $RWT_{IVS+PW} = (IVSth + PWth)/LVDd$; and
RWT_{IVS} = 2 \times \text{IVSth/LVDd}. The American Society of Echocardiography (ASE) recommends RWT_{PW} for calculating RWT [6]. However, some studies found that RWT_{IVS + PW} had clinical significance [7, 8]. The difference in clinical significance among the three methods of measuring RWT is unclear.

To compare the clinical significance of RWT_{PW}, RWT_{IVS + PW}, and RWT_{IVS}, the prognostic values of the RWTs were examined and compared in patients with ADHF.

Materials and methods
Participants
This was a single-center, retrospective, observational study conducted at a Japanese community hospital. In total, 426 consecutive patients admitted due to ADHF through the clinic or emergency room were recruited between June 2014 and April 2016 and followed-up from June 2014 to September 2016. A total of 41 patients were excluded for any of the following reasons: no TTE on admission (n = 35); and RWT not measured (n = 6). Finally, 385 patients were eligible for the analysis (Fig. 1). We previously documented the enrolled patients in detail [5].

The present study followed the tenets of the Declaration of Helsinki and the Ethical Guidelines for Medical and Health Research Involving Human Subjects proposed by the Ministry of Health and Welfare in Japan. The institutional ethics committee at Tomishiro Central Hospital approved the present study and waived informed consent because of the observational nature of the study.

Transthoracic echocardiography
Comprehensive TTE (Vivid 7 ultrasound system, GE Vingmed Ultrasound, Horten, Norway) was performed during hospital admission by four medical technicians who had at least 5 years of experience performing TTE. Their measurements followed established and standardized methods recommended by the ASE and the European Society of Cardiology. At least two attending cardiologists certified by the Japanese Circulation Society and an experienced sonographer reviewed the echocardiography reports immediately after comprehensive TTE. LV geometry, including PWth, IVSth, and LVDd, was measured in M-mode in a parasternal long-axis view [6]. All measurements were performed from the leading edge to the leading edge [6]. RWTs were calculated by the three measurement methods and defined as follows: RWT_{PW} = 2 \times \text{PWth/LVDd}; RWT_{IVS + PW} = (\text{IVSth} + \text{PWth})/\text{LVDd}; and RWT_{IVS} = 2 \times \text{IVSth/LVDd}. The patients were divided into two groups based on the median RWT_{PW} (low- and high-RWT_{PW}), median RWT_{IVS + PW} (low- and high-RWT_{IVS + PW}), or median RWT_{IVS} (low- and high-RWT_{IVS}).

Left ventricular ejection fraction (LVEF) was assessed using the biplane Simpson’s method [6]. Heart failure with preserved ejection fraction (HFpEF) was defined as an ejection fraction ≥50% [9]. LV mass was computed by the Cube formula [6]. LV end-diastolic volume (LVEDVI) was estimated by the Teichholz equation [10]. Peak transmitral early diastolic wave (E wave) velocity, atrial contraction wave (A wave) velocity, and deceleration time (DCT) were measured by the pulse wave Doppler signals of the mitral
inflow in the apical four-chamber view [11]. Valvular diseases were evaluated using a semiquantitative 4-grade scale (none, mild, moderate, and severe) [12].

Data collection
Cardiologists followed the patients at Tomishiro Central Hospital Clinic every 1–3 months after hospital discharge. Medical clerks confirmed the patients’ condition if the patients canceled the appointment. Patients’ medical charts were reviewed to collect their demographic characteristics and clinical data, including medications, laboratory tests, and hemodynamic data on hospital admission. The primary outcome was all-cause death. Death was confirmed by the medical chart, telephone call with a patient’s family, or obituary in local newspapers.

Statistical analysis
Continuous variables with normal and skewed distributions are presented as means (SD) and medians [25th, 75th percentiles], respectively. Categorical variables are presented as numbers with a percentage.

In two-group comparisons, Student’s t-test and the Mann-Whitney U test were used to compare normally distributed and non-normally distributed continuous variables, respectively. Fisher’s exact test was used for categorical variables.

Survival analysis
During follow-up (235 [92, 425] days), 95/385 (25%) patients died. Survival analysis for all-cause death was performed. Kaplan-Meier curves were stratified by RWT PW, RWTIVS + PW, and RWTIVS. The log-rank test was used to compare survival curves. High-RWT PW, high-RWTIVS + PW, and high-RWTIVS were examined by univariate Cox proportional hazard models and a Cox proportional hazard model adjusted by the Get With The Guideline score (GWTG) [13, 14], an established risk score for mortality in patients with acute heart failure, to obtain hazard ratios (HRs) and 95% confidence intervals (95% CIs).

Logistic regression model for 90-day mortality
A total of 48 patients who were lost to follow-up were excluded to evaluate the risk of 90-day mortality. Logistic regression models were used to obtain the odds ratios (ORs) of 90-day mortality and 95% CIs. High-RWT PW, high-RWTIVS + PW, and high-RWTIVS were examined in univariate logistic regression models and a logistic regression model adjusted by GWTG.

Receiver operating curves for 90-day mortality
Receiver operating curves for 90-day mortality were drawn using RWT PW, RWTIVS + PW, and RWTIVS to obtain c-statistics, and the best RWT cut-off values were determined by the maximum Youden index [15].

Sensitivity analysis of the survival analysis by stratified RWTs by the best cut-off
To confirm the consistency of the survival analysis, the participants were divided based on the best RWT cut-off value derived from the Youden index. Survival analysis was performed to compare low and high-RWTs. High-RWT PW, high-RWTIVS + PW, and high-RWTIVS were also examined with univariate and adjusted proportional Cox hazard models.

Relationships between RWTs and clinical characteristics
Spearman’s correlation coefficient (ρ) was used to identify significant associations between RWTs and clinical characteristics: age, the natural logarithm of brain natriuretic peptide (logBNP), LVEF, LVEDV, and systolic blood pressure (SBP).

Reliability of measurement of PWth and IVSth
The reliabilities of the TTE measurements of PWth, IVSth, and LVDd were examined in 25 patients whose TTE image quality was good, and all of the patients underwent TTE performed by the same one of four medical technicians. The medical technician and two other examiners re-measured PWth and IVSth in the TTE image stored in the local server on hospital admission, using an off-line image analysis system (Nahri Aqua, Mehergen Group, Fukuoka, Japan). Comparing every two examiners’ measurements, Bland-Altman plots were used to assess the agreement between the measurement by the same examiner and different examiners [16]. The inter-class coefficient (ICC) was computed to assess agreement [17]. The reliabilities of RWT PW, RWTIVS + PW, and RWTIVS were also examined. RWTs were computed using PWth, IVSth, and LVDd measured by three examiners. Bland-Altman plots were drawn, and the ICC and P values were calculated.

Software
The statistical software used was R 3.4.3 (R Foundation for Statistical Computing, Vienna, Australia). All reported P values are two-tailed, and a P value < 0.05 was considered significant.

Results
Participants
The participants’ median age was 81 years, and there were 181/385 (47%) men in the overall population. Comparing low- and high-RWT PW, high-RWT PW had more elderly patients and more females, whereas in comparisons between low- and high-RWTIVS + PW...
| Demographic data and echocardiographic parameters | Overall | RWT_{PW}^{b} | RWT_{IVS} + PW^{b} | RWT_{IVS}^{b} |
|-----------------------------------------------|--------|-------------|----------------|-------------|
| Age, y | 81 [70, 88] | 80 [69, 87] | 83 [73, 89] | 0.021 | 80 [69, 87] | 83 [73, 89] | 0.067 | 80 [69, 87] | 83 [73, 89] | 0.082 |
| Mele, n (%) | 181/385 (47) | 104/193 (54) | 77/192 (41) | 0.008 | 99/193 (51) | 82/192 (43) | 0.1 | 98/193 (51) | 83/192 (43) | 0.15 |
| Height, cm | 154 ± 10 | 156 ± 10 | 153 ± 10 | 0.002 | 153 ± 10 | 153 ± 10 | 0.084 | 153 ± 10 | 153 ± 10 | 0.15 |
| Body weight, kg | 60 ± 16 | 60 ± 15 | 60 ± 17 | 0.05 | 61 ± 15 | 59 ± 17 | 0.46 | 60 ± 16 | 59 ± 16 | 0.59 |
| Body mass index, kg/m² | 22.8 ± 4.6 | 22.6 ± 4.5 | 23.0 ± 4.8 | 0.42 | 22.8 ± 4.6 | 22.8 ± 4.7 | 0.99 | 22.9 ± 4.7 | 22.8 ± 4.6 | 0.87 |
| Echocardiographic parameters | | | | | | | | |
| RWT_{PW}^{d} | 0.36 ± 0.12 | 0.28 ± 0.05 | 0.45 ± 0.12 | < 0.001 | 0.28 ± 0.05 | 0.44 ± 0.12 | < 0.001 | 0.30 ± 0.07 | 0.43 ± 0.13 | < 0.001 |
| RWT_{IVS} + PW^{d} | 0.37 ± 0.13 | 0.29 ± 0.06 | 0.45 ± 0.12 | < 0.001 | 0.28 ± 0.05 | 0.46 ± 0.12 | < 0.001 | 0.29 ± 0.06 | 0.46 ± 0.12 | < 0.001 |
| RWT_{IVS}^{d} | 0.38 ± 0.14 | 0.30 ± 0.09 | 0.46 ± 0.15 | < 0.001 | 0.29 ± 0.06 | 0.48 ± 0.14 | < 0.001 | 0.28 ± 0.06 | 0.48 ± 0.13 | < 0.001 |
| IVSth, mm | 9.4 ± 2.4 | 8.5 ± 2.0 | 10.4 ± 2.4 | < 0.001 | 8.1 ± 1.8 | 10.7 ± 2.2 | < 0.001 | 7.9 ± 1.5 | 11.0 ± 2.1 | < 0.001 |
| PWth, mm | 9.0 ± 2.1 | 7.5 ± 1.3 | 10.3 ± 2.0 | < 0.001 | 8.1 ± 1.4 | 10.2 ± 2.2 | < 0.001 | 8.3 ± 1.6 | 9.8 ± 2.3 | < 0.001 |
| LVDd, mm | 52 ± 9.7 | 57 ± 8.8 | 46.8 ± 7.8 | < 0.001 | 57.7 ± 8.2 | 46.2 ± 7.6 | < 0.001 | 57 ± 8.8 | 46.8 ± 7.8 | < 0.001 |
| LVEF, (%) | 47 ± 17 | 41 ± 16 | 51 ± 16 | < 0.001 | 39 ± 16 | 52 ± 16 | < 0.001 | 40 ± 16 | 51 ± 16 | < 0.001 |
| HFrEF (LVEF ≥ 50%), n (%) | 157/383 (41) | 56/193 (29) | 101/190 (53) | < 0.001 | 47/193 (24) | 110/190 (58) | < 0.001 | 50/193 (26) | 107/190 (56) | < 0.001 |
| LVM, g | 168 [131, 211] | 173 [135, 211] | 164 [132, 208] | 0.51 | 174 [138, 211] | 164 [133, 208] | 0.18 | 170 [135, 207] | 166 [130, 212] | 0.9 |
| LVDD, mL | 130 [92, 167] | 160 [130, 194] | 102 [79, 126] | < 0.001 | 160 [130, 194] | 97 [74, 124] | < 0.001 | 160 [134, 194] | 102 [79, 130] | < 0.001 |
| LVMI/LVDD | 1.43 ± 0.56 | 1.10 ± 0.23 | 1.75 ± 0.61 | < 0.001 | 1.08 ± 0.20 | 1.77 ± 0.59 | < 0.001 | 1.09 ± 0.21 | 1.76 ± 0.6 | < 0.001 |
| E wave, cm/sec | 97 ± 29 | 97 ± 30 | 97 ± 28 | 0.04 | 95 ± 28 | 99 ± 29 | 0.37 | 98 ± 30 | 96 ± 28 | 0.58 |
| A wave, cm/sec | 76 ± 32 | 70 ± 29 | 82 ± 34 | 0.01 | 73 ± 29 | 79 ± 34 | 0.13 | 74 ± 30 | 77 ± 33 | 0.57 |
| E/A | 1.21 [0.84, 1.83] | 1.29 [0.85, 1.84] | 1.15 [0.82, 1.68] | 0.18 | 1.21 [0.84, 1.84] | 1.21 [0.86, 1.81] | 0.81 | 1.22 [0.84, 1.83] | 1.17 [0.85, 1.83] | 0.6 |
| Deceleration time, ms | 150 [123, 195] | 150 [121, 196] | 150 [128, 192] | 0.04 | 149 [119, 185] | 150 [129, 200] | 0.043 | 147 [118, 181] | 152 [129, 201] | 0.013 |
| Aortic valve stenosis, n (%) | 29/385 (7.5) | 7/193 (3.6) | 22/192 (12) | 0.004 | 8/193 (4.1) | 21/192 (11) | 0.012 | 8/193 (4.1) | 21/192 (11) | 0.012 |
| Aortic valve regurgitation, n (%) | 24/385 (6.2) | 14/193 (7.3) | 10/192 (5.2) | 0.53 | 14/193 (7.3) | 10/192 (5.2) | 0.53 | 11/193 (5.7) | 13/192 (6.8) | 0.68 |
| Table 1 Demographic data and echocardiographic parameters (Continued) |
|-----------------------------------|-----------------|
|                                  | Overall         | RWT<sub>PW</sub><sup>b</sup> | RWT<sub>IVS+PW</sub><sup>b</sup> | RWT<sub>IVS</sub><sup>b</sup> |
|                                  | Low  | High | P value | Low  | High | P value | Low  | High | P value |
|-----------------------------------|------|------|---------|------|------|---------|------|------|---------|
| n = 385                          |      |      |         |      |      |         |      |      |         |
| n = 193                          |      |      |         |      |      |         |      |      |         |
| n = 192                          |      |      |         |      |      |         |      |      |         |
| Mitral valve regurgitation, n (%) | 59/385 (15) | 39/193 (20) | 20/192 (10) | 0.01 | 40/193 (21) | 19/192 (9.9) | 0.004 | 41/193 (21) | 18/192 (9.4) | 0.002 |
| Laboratory data                  |      |      |         |      |      |         |      |      |         |
| Blood urea nitrogen, mg/dL       | 24 [17, 35]  | 24 [17, 34]  | 24 [17, 36]  | 0.77 | 24 [17, 35]  | 24 [17, 35]  | 0.67 | 24 [17, 35]  | 23 [17, 35]  | 0.76 |
| Creatinine, mg/dL                | 1.14 [0.81, 1.52] | 1.15 [0.83, 1.54] | 1.09 [0.79, 1.51] | 0.31 | 1.18 [0.83, 1.56] | 1.09 [0.79, 1.50] | 0.22 | 1.17 [0.83, 1.55] | 1.07 [0.79, 1.50] | 0.19 |
| Hemoglobin, g/dL                 | 12.0 ± 2.4 | 12.0 ± 2.4 | 11.9 ± 2.4 | 0.84 | 11.9 ± 2.4 | 12.0 ± 2.4 | 0.78 | 11.9 ± 2.5 | 12 ± 2.3 | 0.68 |
| Brain natriuretic peptide, pg/mL | 666 [427, 1266] | 737 [449, 1376] | 638 [403, 1155] | 0.056 | 765 [472, 1376] | 636 [401, 1092] | 0.026 | 683 [437, 1349] | 645 [413, 1190] | 0.28 |
| Medication, n (%)                |      |      |         |      |      |         |      |      |         |
| ACE-I and/or ARB                 | 124/285 (32) | 71/193 (37) | 53/192 (28) | 0.069 | 67/193 (35) | 57/192 (30) | 0.34 | 66/193 (34) | 58/192 (30) | 0.47 |
| Beta blocker                     | 153/385 (40) | 78/193 (40) | 75/192 (39) | 0.069 | 75/193 (39) | 78/192 (41) | 0.8 | 74/193 (38) | 79/192 (41) | 0.65 |
| Hemodynamic data                 |      |      |         |      |      |         |      |      |         |
| Systolic blood pressure, mmHg    | 132 ± 26 | 128 ± 24 | 135 ± 29 | 0.006 | 131 ± 24 | 133 ± 29 | 0.35 | 130 ± 24 | 134 ± 29 | 0.13 |
| Diastolic blood pressure, mmHg   | 78 ± 21 | 75 ± 19 | 73 ± 17 | 0.009 | 74 ± 18 | 77 ± 21 | 0.17 | 74 ± 18 | 77 ± 21 | 0.1 |
| Heart rate, bpm                  | 84 ± 21 | 83 ± 21 | 84 ± 21 | 0.81 | 83 ± 19 | 84 ± 22 | 0.64 | 83 ± 19 | 83 ± 22 | 0.78 |

A, late mitral valve inflow velocity; ACE-I, angiotensin converting enzyme inhibitor; ARB, angiotensin receptor blocker; E, early mitral valve inflow velocity; IVS<sub>th</sub>, interventricular septum thickness; LVEDV, left ventricular end diastolic volume; LVDd, left ventricular internal dimension at end-diastole; LVEF, left ventricular ejection fraction; LVM, left ventricular mass; PW<sub>th</sub>, posterior wall thickness; RWT, relative wall thickness

<sup>a</sup>RWT was the ratio of left ventricular wall thickness to LVDd. Left ventricular wall thickness was measured at interventricular septum as IVS<sub>th</sub> and posterior wall as PW<sub>th</sub>. Three measurement methods to compute RWT were as follows: RWT<sub>PW</sub> = 2 × PW<sub>th</sub>/LVDd, RWT<sub>IVS+PW</sub> = (PW<sub>th</sub> + IVS<sub>th</sub>)/LVDd, and RWT<sub>IVS</sub> = 2 × IVS<sub>th</sub>/LVDd

<sup>b</sup>The patients were divided into two groups based on the median of RWT<sub>PW</sub>, RWT<sub>IVS+PW</sub>, and RWT<sub>IVS</sub>.
and between low- and high-RWT<sub>IVS</sub>, there were no significant differences in baseline characteristics (Table 1).

**Transthoracic echocardiography**
The mean RWT<sub>PW</sub>, RWT<sub>IVS + PW</sub>, and RWT<sub>IVS</sub> values in the overall population were 0.36 ± 0.12, 0.37 ± 0.13, and 0.38 ± 0.14, respectively.

On comparing the three RWTs (low- vs. high- RWT<sub>PW</sub>, RWT<sub>IVS + PW</sub>, RWT<sub>IVS</sub>), high-RWTs had thicker IVS<sub>th</sub> and PW<sub>th</sub>, smaller LVD<sub>d</sub>, greater LVEF, smaller LV end-diastolic volume, high LVM/LVEDV, and less severe mitral regurgitation than low-RWTs (Table 1).

**Survival analysis**
During follow-up (235 [92, 425] days), 95/385 (25%) patients died in the overall population.

Comparing low- and high-RWT<sub>PW</sub>, there was a significant difference in the incidence of all-cause death (low 36/193 (19%) vs. high-RWT<sub>PW</sub> 59/192 (31%), P = 0.007). Kaplan-Meier curves showed that high-RWT<sub>PW</sub> had worse survival than low-RWT<sub>PW</sub> (P for log-rank test = 0.009; Fig. 2a).

Comparing low- and high-RWT<sub>IVS + PW</sub>, there was no significant difference in all-cause death (low 40/193 (21%) vs. high-RWT<sub>PW</sub> 55/192 (29%), P = 0.077) or survival (P for log-rank test = 0.074; Fig. 2b).

In a comparison between low- and high-RWT<sub>IVS</sub>, there was no significant difference in all-cause death (low 42/193 (22%) vs. high-RWT<sub>IVS</sub> 53/192 (28%), P incidence = 0.2) or survival (P for log-rank test = 0.19; Fig. 2c).

**Cox proportional hazard models for all-cause death**
In the unadjusted and adjusted Cox proportional hazard models, high-RWT<sub>PW</sub> was a significant risk factor for all-cause death (unadjusted Cox model, HR (95% CI), 1.72 (1.41–2.17), P = 0.01; adjusted Cox model, 1.95 (1.28–2.98), P = 0.02; Table 2).

High-RWT<sub>IVS + PW</sub> was not a significant risk factor for all-cause death in the unadjusted Cox proportional model (unadjusted Cox model, HR, 1.45 (0.96–2.17), P = 0.075), but it was in the adjusted Cox proportional hazard model (adjusted Cox model, 1.53 (1.01–2.32), P = 0.045; Table 2).

High-RWT<sub>IVS</sub> was not a significant factor in either the unadjusted or the adjusted Cox proportional hazard model (Table 2).

**Logistic regression models for 90-day mortality**
The OR of high- to low-RWT<sub>PW</sub> was significant (univariate, OR, 2.19, 95%CI, 1.15–2.19, P = 0.017; adjusted, OR, 2.26, 95%CI, 1.16–4.4, P = 0.017) on univariate analysis and the adjusted logistic regression model (Table 3). In contrast, the OR of neither high to low-RWT<sub>IVS + PW</sub> nor RWT<sub>IVS</sub> was significant on univariate analysis or the adjusted logistic regression models.

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Table 2 Cox proportional hazard model for evaluate the risk of RWTs for all-cause mortality

| Calculate method and factor | Unadjusted | Adjusted by GWTG |
|----------------------------|------------|------------------|
| Event/cases                | Event/cases* |
| HR                         | HR         |
| 95% CI                     | 95% CI     |
| P value                    | P value    |

|                  | Unadjusted | Adjusted by GWTG |
|------------------|------------|------------------|
| High-to low-RWT<sub>PW</sub> | 95/385 | 93/380 |
| 1.72             | 1.95       |
| 1.14             | 1.28       |
| 0.01             | 2.98       |
| High-to low-RWT<sub>IVS + PW</sub> | 95/385 | 93/380 |
| 1.45             | 1.53       |
| 0.96             | 1.01       |
| 0.075            | 2.32       |
| High-to low-RWT<sub>IVS</sub> | 95/385 | 93/380 |
| 1.31             | 1.36       |
| 0.87             | 0.9        |
| 0.19             | 2.06       |

CI confidence interval; GWTG Get With The Guideline score; HR hazard ratio; RWT relative wall thickness

*S cases were removed because of GWTG missing
Receiver operating curves for 90-day mortality

A total of 48 (13%) patients died within 90 days from hospital admission. Figure 3 shows the receiver operating characteristic (ROC) curves for 90-day mortality using the RWTs. The c-statistic of the ROC curve using RWT PW was 62.6%, and the best cut-off value was 0.35. The c-statistic of the ROC curve using RWTIVS + PW was 59.7%, and the best cut-off value was 0.55. The c-statistic of the ROC curve using RWT IVS was 43.1%, and the best cut-off value was 0.36.

Sensitivity analysis of the survival analysis by stratified RWTs by the best cut-off

Additional file 1: Table S1 shows the demographic data and echocardiographic data with stratification by the best RWT cut-off. High-RWT PW had worse survival than low-RWT PW (P for log-rank test = 0.03; Additional file 2: Figure S1a). High-RWTIVS + PW also had a worse prognosis than low-RWT IVS + PW (P for log-rank test < 0.001; Additional file 2: Figure S1b). In contrast, there was no significant difference in survival between low- and high-RWT IVS (P for log-rank test = 0.077; Additional file 2: Figure S1c).

In the unadjusted and adjusted Cox proportional hazard models, high-RWT PW and high-RWTIVS + PW were associated with mortality (high-RWT PW, unadjusted Cox model, HR (95% CI), 1.55 (1.04–2.33), P = 0.033; adjusted Cox model, 1.72 (1.14–2.59), P = 0.01; high-RWTIVS + PW, unadjusted Cox model, HR (95% CI), 3.88 (2.34–6.43), P < 0.001; adjusted Cox model, 3.42 (2.04–5.72), P < 0.001; Additional file 3: Table S2). High-RWT IVS was not a significant risk factor in the unadjusted and adjusted Cox proportional hazard models.

Relationship between RWTs and clinical characteristics

There were significant positive correlations between the three RWTs and age and LVEF, and negative correlations between the RWTs and LogBNP and LVEDV (Table 4). RWTIVS + PW and RWTIVS did not have significant correlations with systolic blood pressure, but RWT PW did (ρ = 0.15, P = 0.004).

Reliability of TTE measurement of PWth, IVSth, and LVDd

Intra-observer agreement of TTE measurement of PWth was significant (ICC = 0.73, P < 0.001; Fig. 4a). Inter-observer agreements of TTE measurement of PWth were also significant (observer 1 vs. 2, ICC = 0.76, P < 0.001; observer 1 vs. 3, ICC = 0.6, P < 0.001; observer 2 vs. 3, ICC = 0.7, P < 0.001; Fig. 3a). There were no systematic biases in the intra- and inter-observer agreements in PWth measurement (Fig. 4a).

Intra-observer agreement of TTE measurement of IVSth was significant (ICC = 0.88, P < 0.001; Fig. 4b). Inter-observer agreements of TTE measurement of IVSth were also significant (observer 1 vs. 2, ICC = 0.81, P < 0.001; observer 1 vs. 3, ICC = 0.77, P < 0.001; observer 2 vs. 3, ICC = 0.73, P < 0.001; Fig. 4b). There were no systematic biases in the intra- and inter-observer agreements in IVSth measurement (Fig. 4b).

Intra-observer agreement of TTE measurement of LVDd was significant (ICC = 0.94, P < 0.001; Fig. 4c). Inter-observer agreements of TTE measurement of LVDd were also significant (observer 1 vs. 2, ICC = 0.71,
Table 4 Relationship between RWTs and clinical characteristics

|                      | RWT_{PW}     |                      | RWT_{IVS+PW}  |                      | RWT_{IVS}     |
|----------------------|--------------|----------------------|---------------|----------------------|--------------|
|                      | p            | P value              | p             | P value              | p            | P value              |
| Age, y               | 0.15         | 0.003                | 0.17          | 0.003                | 0.17         | 0.001                |
| LogBNP, log (pg/mL)  | −0.2         | < 0.001              | −0.15         | 0.003                | −0.11        | 0.039                |
| LVEF, %              | 0.42         | < 0.001              | 0.47          | < 0.001              | 0.43         | < 0.001              |
| LVEDV, mL            | −0.67        | < 0.001              | −0.74         | < 0.001              | −0.69        | < 0.001              |
| Systolic blood pressure, mmHg | 0.15 | 0.004 | 0.094 | 0.065 | 0.063 | 0.22 |

LogBNP logarithmed brain natriuretic peptide; LVEDV left ventricular end-diastolic volume; LVEF left ventricular ejection fraction; p, Spearman’s correlation coefficient

Intra-observer agreement of RWT_{PW} was significant (ICC = 0.77, P < 0.001; Fig. 5a). Inter-observer agreements of RWT_{PW} were also significant (observer 1 vs. 2, ICC = 0.74, P < 0.001; observer 1 vs. 3, ICC = 0.65, P < 0.001; observer 2 vs. 3, ICC = 0.63, P < 0.001). There were no systematic biases in the intra- and inter-observer agreements in RWT_{PW}.

Intra-observer agreement of RWT_{IVS+PW} was significant (ICC = 0.89, P < 0.001; Fig. 5b). Inter-observer agreements of RWT_{PW} were also significant (observer 1 vs. 2, ICC = 0.82, P < 0.001; observer 1 vs. 3, ICC = 0.74, P < 0.001; observer 2 vs. 3, ICC = 0.83, P < 0.001). There were no systematic biases in the intra- and inter-observer agreements in RWT_{IVS+PW}.

Intra-observer agreement of RWT_{IVS} was significant (ICC = 0.84, P < 0.001; Fig. 5c). Inter-observer agreements of RWT_{IVS} were also significant (observer 1 vs. 2, ICC = 0.77, P < 0.001; observer 1 vs. 3, ICC = 0.75, P < 0.001; observer 2 vs. 3, ICC = 0.72, P < 0.001). There were no systematic biases in the intra- and inter-observer agreements in RWT_{IVS}.
Discussion

To the best of our knowledge, this is the first study to show the difference in the clinical significance of the three RWTs. The present study demonstrated that, compared to RWT_{IVS + PW} and RWT_{IVS}, RWT_{PW} is the best to stratify the risk for all-cause death in ADHF patients. This may be consistently supported by three findings. First, high-RWT PW had a significantly worse prognosis than low-RWT PW. In contrast, on survival analysis, there was no significant difference between high- and low-RWT_{IVS + PW} or RWT_{IVS}. Second, in the logistic regression model for 90-day mortality, only high-RWT PW was significant among the three RWTs (Table 3). Third, ROC for 90-day all-cause death using RWT PW had the highest c-statistic among the three ROCs.

Explanations of the differences in the prognostic values among the three RWTs

High-RWT_{PW} was associated with a poor prognosis on survival analysis and Cox proportional hazard models (Fig. 2a; Table 2). High-RWT_{IVS + PW} was not associated with poor survival on survival analysis (Fig. 1b), whereas high-RWT_{IVS + PW} was a significant risk only in the Cox proportional hazard model adjusted by GWTG, not in the unadjusted model (Table 2). High-RWT_{IVS} did not show worse survival than low-RWT_{IVS} (Fig. 1c; Table 2).

The equations of RWT_{PW} and RWT_{IVS + PW} contain PWth. PWth or the ratio of PWth to LVDd, therefore, may represent the LV remodeling related to a worse prognosis better than IVSth or IVSth to LVDd in patients with ADHF. Patients with high-RWT_{PW} had higher systolic blood pressure than those with low-RWT_{PW} (Table 1), while there was no such difference between high-RWT_{IVS + PW} or between low- and high-RWT_{IVS}. RWT_{PW} had a positive correlation with systolic blood pressure (Table 4), while either RWT_{IVS + PW} or RWT_{IVS} did not. This may suggest that thickening of PWth, rather than IVSth, is likely to counterbalance pressure overload and may lead to LV diastolic dysfunction leading to a poor prognosis. A higher A wave in high RWT_{PW} patients than in low RWT_{PW} patients may support this assumption (Table 1).

In terms of methodological validity, there were no differences in inter- and intra-observer agreements for each RWT. Given that fairly good reproducibility was observed in all measurements, differences in prognostic values among the three RWTs may not result from technical aspects of TTE.

Paradoxically, high-RWT_{PW} patients had lower BNP than low-RWT_{PW} patients (Table 1). High-RWT_{PW} included 101 (53%) patients with HfPeF. Generally, BNP increases modestly in HfPeF [18]. Furthermore,
the prognostic value of BNP has not been confirmed in patients with HFpEF [19]. High RWT_{PW} might be of clinically utility, especially, in patients with HFpEF.

Limitations
The present study had several limitations. The present study did not have pressure data such as LV end-diastolic pressure or pulmonary artery wedge pressure. LV wall thickness was not evaluated by other modalities, such as magnetic resonance imaging or computed tomography. Patients having valvular diseases with various etiologies were included, which might affect the prognostic value of RWTs.

In conclusion, high-RWT_{PW} had a higher systolic pressure and A wave than low-RWT_{PW}. This finding was not observed in the comparison between low- and high-RWT_{IVS-PW} or between low-and high-RWT_{IVS-PW}th may represent pressure overload better than IVSth. When calculating RWT, RWT_{PW} should be recommended for evaluating the mortality risk in ADHF.

Supplementary information
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Authors’ contributions
SY originated the concept of this study, collected data, and wrote the original manuscript. MA and TA collected the data and discussed the study results. OA took responsibility for conducting the study. MS and SU supervised the study concept and discussed the study results in depth. All authors approved the final manuscript submission.

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The institutional ethics committee at Tomishiro Central Hospital approved the present study and waived informed consent because of the observational nature of the study.

Consent for publication
Not applicable.

Competing interests
The authors declare that they have no competing interests.

Author details
1Department of Cardiology, Tomishiro Central Hospital, 25 Ueta, Okinawa 901-0243, Japan. 2Department of Clinical Pharmacology and Therapeutics, Graduate School of Medicine, University of the Ryukyus, 205 Uehara, Nishihara-cho, Okinawa 901-0215, Japan. 3Department of Diabetes, Endocrinology, and Metabolism, School of Medicine, Fukushima Medical University, 1 Hikarigaoka, Fukushima 960-1295, Japan.

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