RESEARCH ARTICLE

General practitioners’ consultation counts and associated factors in Swiss primary care – A retrospective observational study

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Abstract

Background

Research on individual general practitioner (GP) workload, e.g. in terms of consultation counts, is scarce. Accurate measures are desirable because GPs’ consultation counts might be related to their work satisfaction and arguably, there is a limit to the number of consultations a GP can hold per day without jeopardizing quality of care. Moreover, understanding the association of consultation counts with GP characteristics is crucial given current trends in general practice, such as the increasing proportion of female GPs, part-time work and group practices.

Aim

The aim of this study was to describe GPs’ consultation counts and efficiency and to assess associations with GP and practice variables.

Methods

In this retrospective observational study we used routine data in electronic medical records obtained from 245 Swiss GPs in 2018. We described GPs’ daily consultation counts as well as their efficiencies (i.e. total consultation counts adjusted for part-time work) and used hierarchical linear models to find associations of the GPs’ total consultation counts in 2018 with GP- and practice-level variables.

Results

The median daily consultation count was 28 over all GPs and 33 for full-time working GPs. Total consultation counts increased non-linearly with part-time status, with high part-time working GPs (60%-90% of full-time) being equally or more efficient than full-time workers. Excluding part-time status in the regression resulted in higher consultation counts for male GPs working in single practices and with older patients, whereas part-time adjusted consultation counts were unaffected by GP gender and practice type.
Conclusion
Female gender, part-time work in the range of 60%-90% of full-time, and working in group practices do not decrease GP efficiency. However, the challenge of recruiting sufficient numbers of GPs remains.

Introduction
Research on individual general practitioner (GP) workload, e.g. in terms of consultation counts, is scarce. So far, studies of consultation counts have depended on self-reports which are highly susceptible to recall bias and disregard day-to-day variability and often also part-time status [1–3]. More accurate measures are desirable because GPs’ consultation counts might be related to their work satisfaction [4] and arguably, there is a limit to the number of consultations a GP can hold per day without jeopardizing quality of care.

Moreover, current trends suggest an increasing demand for primary care consultations [5–7]. The demand is likely to increase further because of accumulating chronic conditions in the aging population [8–12]. At the same time, the GP population is changing. In Switzerland and other occidental countries, many GPs will reach retirement age soon [3, 13]. The next generation of GPs will consist of an increased proportion of female individuals and will prefer working part-time and in urban group practices [1, 3, 14–17]. Whether these next-generation GPs with different preferences will be able to fill the gap of the retiring ones is uncertain as, until now, little is known about the associations between GPs’ personal characteristics and their consultation output.

Therefore, the aim of this study was to describe GPs’ consultation counts and efficiency and to assess associations with GP and practice variables using electronic medical records (EMR) data.

Methods
2.1 Design and setting
We conducted a retrospective observational study with data obtained from the FIRE (Family Medicine ICPC-Research using Electronic Medical Records) database. The FIRE database collects anonymized routine data exported from EMR of participating GPs from the German speaking part of Switzerland. Since the project started in 2009, 524 GPs (roughly 10% of all GPs working in the German speaking area [18]) have joined. The database holds records of over 623’000 patients and more than 6.9 million consultations (as of April 04, 2019). For this cross-sectional study, we restricted our analyses to consultations on workdays (see definition below) in 2018. According to the Ethics Committee of the Canton of Zurich, the project does not fall under the scope of the Federal Act on Research involving Human Beings (Human Research Act) [19] and therefore no ethical consent was necessary (BASEC-Nr: Req-2017-00797).

2.2 Participants
We included all GPs in the FIRE database with a) known age and part-time status (% full-time equivalent) in 2018. We excluded those who b) exported data of less than 10 months in 2018, c) were associated with multiple GP practices (e.g. because of changing their workplace from one practice to another), d) exported data as a group of GPs (precluding analyses of individual...
GPs' consultation counts), e) showed evidence for false-negative consultation data (incomplete datasets) and f) belonged to part-time working strata containing < 5 GPs (disallowing meaningful summary statistics).

To improve the accuracy of the inclusion criterion and data validity, we updated the part-time status by email inquiry to all participating practices. In this way, we achieved an inclusion rate of 88% (312 of 353) of individual FIRE GPs in 2018. To improve the accuracy of the exclusion criteria, we examined the plausibility of consultation counts by manually investigating outlier GPs, defined as those with part-time adjusted total consultation counts in 2018 below the 10th or above the 90th percentile. Outlier GPs were investigated by searching their practice websites and additional internal information for other GPs that exported data under the same GP identifier (leading to exclusion under criterion d) and by checking whether they did not export properly (leading to exclusion under criterion e). The selection process is visualized in the study flowchart (Fig 1).

2.3 Database query, variables and definitions

From the database, we extracted consultation data for all workdays in 2018. As workdays we considered all days of the year except weekends and public holidays. Consultation data included patient information (patient identification number, age, and gender), GP information (GP identification number, part-time status, age, gender, employment contract) and practice information (practice identification number, zip code, practice type). Urbanity of the practice was determined from the zip code [20].

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Fig 1. Flowchart.

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GPs included (n=312)
- practices (n=143)
- patients (n=292'304)
- consultations (n=1'566'720)

GPs excluded (n=67)
- export covered less than 10 months (n=47)
- GP changed practice (n=1)
- export as group of GPs (n=14)
- incomplete export (n=4)
- part-time status stratum < 5 GPs (n=1)

GPs analyzed (n=245)
- practices (n=113)
- patients (n=242'551)
- consultations (n=1'285'928)
2.4 Outcomes
Outcomes of the study were:
1. GPs’ daily consultation counts, stratified by part-time status
2. GPs’ efficiencies (consultation counts per full-time equivalent)
3. Associations of GPs’ total consultation counts in 2018 with GP-level and practice-level covariates

2.5 Data analysis
We described categorical data by counts and/or proportions (n, %) as appropriate and numerical data by median and interquartile range (IQR). We aggregated data to represent consultation counts for every GP and workday in 2018. From that, we determined the mean daily consultation count for every GP considering only days when the individual GP held at least one consultation. GPs’ efficiencies, i.e. consultation counts per full-time equivalent, were calculated from the GPs’ total consultation counts in 2018 and their part-time status. They were reported as relative efficiencies with respect to the median number of consultations of full-time working GPs in 2018.

We used hierarchical linear models with random practice effects to find associations of the GPs’ total 2018 consultation counts with GP-level covariates (part-time status, gender, age group, employment contract, and consultation patient characteristics; the latter meaning median age of patients in consultations and proportion of consultations with female patients) and practice-level covariates (practice type, urbanity). Firstly, every variable was included as the only fixed effect amongst the random practice effects (crude model). Secondly, we adjusted using two multivariable models: a model adjusting for all variables except for part-time status, and a fully adjusted model. The rationale behind this was the assumption that variables potentially affect GP consultation counts both indirectly through their effects on part-time status and directly. Disregarding part-time status can therefore give hints on total effects, whereas adjusting for it blocks mediation, thus revealing direct effects. Significance was determined at the 5% level; 95% confidence intervals (CI) were reported accordingly. Variables were adapted for analysis as appropriate, i.e. GP age was categorized into age groups of 10 years and part-time status was rounded to the smaller tens digit. Consultation patient characteristics were calculated for each GP from all their consultations (thus allowing to count individual patients multiple times in order to reflect the GP perspective on consultations) and mean centered. All analyses were conducted using the R statistical package version 3.5.0 [21].

Results
Population
Data covered 1’285’928 consultations with 242’551 individual patients and 245 GPs in 113 practices. Of the GPs, 38.0% were female and the GPs’ median age was 51 (IQR = 43 to 58). The majority worked in group practices (87.3%), was self-employed (64.7%) and located in urban areas (75.5%). On average (median), the GPs worked on 88.8% (IQR = 75.5% to 97.6%) of all workdays and held 4’843 (IQR = 3’318 to 6’908) consultations in 2018 with 1’125 (IQR = 836 to 1’477) different patients. In those consultations, on average (median over GPs), 52.1% (IQR = 49.2% to 59.3%) of patients were female and their median age was 58 years (IQR = 52 to 64). Characteristics of GPs stratified by part-time status are given in Table 1.
Consultation counts per day

The median of the GPs’ mean daily consultation counts was 28 (IQR = 22 to 35). Fig 2 depicts the distribution of daily consultation counts stratified by part-time status. Median daily consultation counts increased with part-time status, revealing three steps with similar values (for 40% to 50%, 60% to 70%, and 80% to full-time, respectively).

Efficiency

GPs’ efficiency, based on part-time adjusted total GP consultation counts in 2018, varied with part-time status. Low (30%-50%) part-time workers were less efficient than full-time workers, whereas high (60% to 90%) part-time workers were equal or more efficient than full-time workers (Fig 3).

Associations of GP- and practice-level factors with total consultation counts

Crude models. The GPs’ total consultation counts in 2018 were highly associated with part-time status in the crude model. All GPs except for those with 90% part-time status held fewer consultations than GPs working full-time. Furthermore, consultation counts were lower among female GPs (-29.9% consultations), employed GPs (-21.3%), and GPs with high proportions of consultations by female patients (-1.7% per one percent increase in female patients). In contrast, consultation counts were higher among GPs working in single practices.
(+37.8%) and GPs with older patients (+1.3% per one year increase in median patient age).

The detailed results of the crude analyses are shown in Table 2.

Adjusted without part-time status. When adjusting for all variables simultaneously except for part-time status, consultation counts were still lower among female GPs (-19.9%) and higher for GPs working in single practices (+18.8%) and with older patients (+0.7% per one year increase in median patient age, Table 3). Employment status and patient gender were no longer significantly associated with total consultation counts. Instead, the latter were now negatively associated with GP age group 60–69 years (-14.7% with respect to age group 50–59 years).

Adjusted including part-time status. Including part-time status in addition to all other predictors dissolved most associations of GP characteristics with total consultation counts. Only oldest GP age (60–69 years) still showed a negative effect (-10.6% with respect to age group 50–59 years), while old patient age showed a positive effect (+0.4% per one year increase in median patient age, Table 3). Part-time status below 90% was still associated with lower consultation counts, but effect sizes were smaller than in the crude model.
Discussion

The median of the GPs’ mean daily consultation counts was 28 over all GPs, and 33 for full-time workers. Daily consultation counts were non-linearly dependent on part-time status; a plateau was reached at 80% part-time status. High part-time working GPs (60%-90% of full-time) were slightly more efficient than full-time workers, with 90% part-time workers having the same total consultation count as full-time workers. Crude associations alone might suggest that highest consultation counts can be found among male GPs working self-employed in single practices caring for predominantly elderly male patients. However, the multilevel regression models put this into perspective: When adjusting for all variables except for part-time status, the effect of GP gender was reduced, and employment status and patient gender were no longer associated with consultation counts. After additionally including part-time status in the model, apart from part-time status itself, only GP and patient age remained significant predictors of consultation counts.

Comparison with existing literature

Daily consultation counts as found in our study were similar to results of a European survey conducted in 1993, both over all GPs and adjusted for part-time work [2]. However, for
Switzerland, more recent studies have reported 24–25 consultations per day [1, 22]. Lower consultation counts in Switzerland compared to other European countries are plausible because consultation count has been shown to be inversely related to consultation duration [23], which is known to be longer in Switzerland [24]—arguably due to the Swiss payment system which considers consultation length for remuneration [1]. The discrepancies between our results and figures from previous Swiss studies may be explained by the different types of consultations considered. While previous surveys inquired only face-to-face contacts, we

| Variable                | Change in consultation count (n) | 95% CI            | p-value   |
|-------------------------|---------------------------------|-------------------|-----------|
| Part-time status        |                                 |                   |           |
| (ref. full-time)        |                                 |                   |           |
| Intercept               | 6924                            | 6444 to 7404      | <0.001    |
| 90%                     | -331                            | -1232 to 570      | 0.471     |
| 80%                     | -944                            | -1537 to -351     | 0.002     |
| 70%                     | -1516                           | -2284 to -749     | <0.001    |
| 60%                     | -2432                           | -3066 to -1797    | <0.001    |
| 50%                     | -3651                           | -4264 to -3038    | <0.001    |
| 40%                     | -4247                           | -5063 to -3431    | <0.001    |
| 30%                     | -4647                           | -5873 to -3420    | <0.001    |
| GP gender               |                                 |                   |           |
| (ref. male)             |                                 |                   |           |
| Intercept               | 6260                            | 5816 to 6705      | <0.001    |
| Female                  | -1873                           | -2347 to -1399    | <0.001    |
| GP age group            |                                 |                   |           |
| (ref. 50–59 years)      |                                 |                   |           |
| Intercept               | 6009                            | 5427 to 6591      | <0.001    |
| 30–39 years             | -879                            | -1761 to 3        | 0.051     |
| 40–49 years             | -601                            | -1288 to 86       | 0.087     |
| 60–69 years             | -632                            | -1465 to 201      | 0.137     |
| Employment status       |                                 |                   |           |
| (ref. self-employed)    |                                 |                   |           |
| Intercept               | 5943                            | 5463 to 6424      | <0.001    |
| Employed                | -1265                           | -2050 to -480     | 0.002     |
| Practice type           |                                 |                   |           |
| (ref. group practice)   |                                 |                   |           |
| Intercept               | 5153                            | 4672 to 5634      | <0.001    |
| Single practice         | 1947                            | 936 to 2958       | <0.001    |
| Urbanity                |                                 |                   |           |
| (ref. urban)            |                                 |                   |           |
| Intercept               | 5309                            | 4774 to 5843      | <0.001    |
| Non-urban               | 946                             | -9 to 1902        | 0.052     |
| Cons. patient characteristics*: gender |               |                   |           |
| Intercept               | 5564                            | 5137 to 5991      | <0.001    |
| % female                | -95                             | -122 to -68       | <0.001    |
| Cons. patient characteristics*: age |               |                   |           |
| Intercept               | 5488                            | 5043 to 5932      | <0.001    |
| median age              | 73                              | 43 to 102         | <0.001    |

Abbreviations: ref. = reference; cons. = consultation; CI = confidence interval

* For continuous predictor variables, coefficients represent the change in consultation count per one unit change.

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considered all types of patient care that led to entries in EMR, including telephone consultations and record reviews.

The association of consultation counts with part-time status was non-linear for both daily and total consultation counts. Daily consultation counts considering only days when the GP actually worked are not representative of the GPs’ outputs but rather of their actual working patterns, e.g. full vs. half days. Interestingly, there seemed to be no difference in daily consultation counts between 40% and 50% part-time workers, between 60% and 70% part-time workers, and among the above 80% part-time workers, respectively, so the difference in total consultation counts must have resulted from the difference in days off (see also Table 1). Total consultation counts in 2018 represent the consultation output irrespective of workday patterns and are thus an appropriate basis for efficiency calculations. We found that high part-time workers (60%-90% of full-time) had a higher efficiency than full-time workers. Though other authors have not stratified part-time status into several categories, they nevertheless observed higher productivity for part-time GPs [25].

Table 3. Multivariable analyses of total consultation counts.

| Variables                              | without part-time status | including part-time status |
|----------------------------------------|--------------------------|---------------------------|
|                                        | Change in consultation count (n) | 95% CI | p-value | Change in consultation count (n) | 95% CI | p-value |
| Intercept                              | 5994                     | 5251 to 6738              | <0.001 | 6734                     | 5982 to 7486  | <0.001 |
| Part-time status (ref. full-time)      |                          |                         |        |                          |                         |        |
| 90%                                    | -                        | -                        | 87     | -1129 to 954              | 0.869 |
| 80%                                    | -                        | -                        | -690   | -1362 to -18              | 0.044 |
| 70%                                    | -                        | -                        | -1249  | -2138 to -360             | 0.006 |
| 60%                                    | -                        | -                        | -2227  | -3011 to -1444            | <0.001 |
| 50%                                    | -                        | -                        | -3127  | -3990 to -2263            | <0.001 |
| 40%                                    | -                        | -                        | -3591  | -4678 to -2504            | <0.001 |
| 30%                                    | -                        | -                        | -4118  | -5554 to -2681            | <0.001 |
| GP gender (ref. male)                  | -1194                    | -2016 to -373             | 0.004  | -162                     | -879 to 555    | 0.658  |
| female                                 |                          |                         |        |                          |                         |        |
| GP age group (ref. 50–59 years)        |                          |                         |        |                          |                         |        |
| 30–39 years                            | -206                     | -1118 to 706              | 0.658  | -308                     | -1072 to 455    | 0.428  |
| 40–49 years                            | -229                     | -896 to 438               | 0.501  | -14                      | -586 to 558     | 0.961  |
| 60–69 years                            | -879                     | -1640 to -119             | 0.023  | -716                     | -1371 to -61    | 0.032  |
| Employment status (ref. self-employed) | -594                     | -1349 to 161              | 0.123  | -200                     | -863 to 463     | 0.554  |
| employed                               |                          |                         |        |                          |                         |        |
| Practice type (ref. group practice)    | 1128                     | 111 to 2144               | 0.030  | 350                      | -622 to 1322    | 0.480  |
| Single practice                        |                          |                         |        |                          |                         |        |
| Urbanity (ref. urban)                  | 327                      | -582 to 1235              | 0.481  | 404                      | -437 to 1245    | 0.346  |
| Non-urban                              |                          |                         |        |                          |                         |        |
| Cons. patient characteristics*:        |                          |                         |        |                          |                         |        |
| % female                               | -40                      | -83 to 4                  | 0.075  | -4                       | -42 to 34       | 0.826  |
| median age                             | 39                       | 7 to 70                   | 0.015  | 27                       | 0 to 54         | 0.049  |

Abbreviations: ref. = reference; cons. = consultation; CI = confidence interval
* For continuous predictor variables, coefficients represent the change in consultation count per one unit change.

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As consultation workload was heavily influenced by part-time status, any other variable’s association with consultation counts must depend on the variable’s own relation with part-time status. Investigations of GP part-time status typically focus on gender differences. In many occidental countries, including Switzerland, female GPs have been found to work part-time, or declare that they plan to do so, more often than male GPs [1, 17, 26–32]. However, many of these studies date back several years and it is hypothesized that work-life-balance choices leading to part-time work might be an issue of ongoing societal change. Therefore, these gender effects could diminish in the future [7, 33, 34]. Today, our crude analysis still revealed a 30% lower crude consultation count for female GPs compared to their male peers. The disappearance of this association after adjustment for part-time working is in line with previous European studies [2, 33]. Interestingly, the association was reduced by one third even in the adjusted model where part-time work was not taken into account, indicating that other variables—such as the GPs’ age, practice type, or characteristics of their patient base—co-transmit the effect. Fittingly, our analyses (consistently with the literature [32, 33]) revealed that GPs aged 50–59 years and those who work in single practices—groups where female GPs are underrepresented [2]—held more consultations. Additionally, female GPs have been reported to care for a higher proportion of female and younger patients [2], which was negatively associated with consultation counts in the crude model. Therefore, part of the gender differences in consultation counts can be explained plausibly by different work settings and patient populations.

**Strengths and limitations**

To our knowledge, this is the first study of this scale using routine data for a detailed investigation of GP consultation counts. We used a large dataset, containing over one million consultations generated by 245 GPs. The inclusion of part-time status was crucial to give insight into consultation workload of individual GPs, given that part-time work has become increasingly common. The combination of multiple regression models allowed for exploration of direct and indirect effects of the investigated variables on consultation counts.

Our GP sample is representative for the Swiss GP community in terms of gender and part-time status but slightly over-represents younger GPs, GPs working employed and those in group practices in urban and suburban areas [17]. Given that future GPs will tend towards working in such environments, our GP sample may better represent the future workforce. Since using EMR is required for participation in the FIRE project, GPs still operating with paper-based medical records were excluded. This part of the workforce, however, can be expected to become less relevant in the future.

The small number of GPs with very low part-time status caused imprecise estimates, therefore the results within these subgroups should be considered with caution. Further subgrouping of the non-urban GP population in order to e.g. analyze the workload of GPs in rural environments was not possible because the sample sizes were too small. Ultimately, our study disregarded consultations on weekends and public holidays and we expect the true total consultation counts in 2018 to be slightly higher. Excluding weekends and public holidays, however, was necessary because on such days, between-GP as well as within-GP variation of consultation counts was very strong and incompatible with our study aim to model typical consultation counts of Swiss GPs.

**Implications for practice and policy**

Knowledge about GPs’ consultation counts can contribute to health policy and health economical decisions [7, 14, 35]. Very high consultation counts may be used as indicators for
compromised quality of care and GP work satisfaction, while very low consultation counts might touch on the healthcare systems’ financial sustainability and raise concerns about the security of future primary care supply.

Our findings suggest that 60%-90% part-time working GPs are at least as efficient as full-time GPs and that efficiency does neither depend on GP gender, employment status nor practice type. Nevertheless, with part-time work becoming more common, the challenge of recruiting new GPs to secure the future workforce remains.

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References
1. Cohidon C, Cornuz J, Senn N. Primary care in Switzerland: evolution of physicians’ profile and activities in twenty years (1993–2012). BMC Family Practice. 2015; 16(1). https://doi.org/10.1186/s12875-015-0321-y PMID: 26292762
2. Boerma WG, van den Brink-Muinen A. Gender-related differences in the organization and provision of services among general practitioners in Europe: a signal to health care planners. Medical care. 2000; 38(10):993–1002. Epub 2000/10/06. https://doi.org/10.1097/00005650-200010000-00003 PMID: 11021672.
3. Merlo P. Work Force Hausarztmedizin in der Schweiz 2015: University of Basel; 2015.
4. Merçay C. Médecins de premier recours–Situation en Suisse, tendances récentes et comparaison internationale. Neuchâtel: Observatoire suisse de la santé (Obsan), 2015.
5. Hippisley-Cox J, Vinogradova Y. Trends in Consultation Rates in General Practice 1995/1996 to 2008/2009: Analysis of the QResearch® database. QResearch®, University of Nottingham and The Health and Social Care Information Centre, 2009.
6. Hobbs FDR, Bankhead C, Mukhtar T, Stevens S, Perera-Salazar R, Holt T, et al. Clinical workload in UK primary care: a retrospective analysis of 100 million consultations in England, 2007–14. The Lancet. 2016; 387(10035):2323–30. https://doi.org/10.1016/S0140-6736(16)00620-6 PMID: 27059886
7. Baird B, Charles A, Honeyman M, Maguire D, Das P. Understanding pressures in general practice. London: The King’s Fund; 2016.
8. Cassell A, Edwards D, Harshfield A, Rhodes K, Brimicombe J, Payne R, et al. The epidemiology of multimorbidity in primary care: a retrospective cohort study. Br J Gen Pract. 2018; 68(669):e245–e51. Epub 2018/03/14. https://doi.org/10.3399/bjgp18X695465 PMID: 29530918; PubMed Central PMCID: PMC5863678.
9. van Oostrom SH, Picavet HSJ, de Bruin SR, Stirbu I, Korevaar JC, Schellevis FG, et al. Multimorbidity of chronic diseases and health care utilization in general practice. BMC Family Practice. 2014; 15(1):61. https://doi.org/10.1186/1471-2296-15-61 PMID: 24708798
10. Harrison C, Henderson J, Miller G, Britt H. Predicting patient use of general practice services in Australia: models developed using national cross-sectional survey data. BMC Family Practice. 2019; 20(1):28. Epub 2019/02/16. https://doi.org/10.1186/s12875-019-0914-y PMID: 30764778; PubMed Central PMCID: PMC6376650.

11. Mukhtar TK, Bankhead C, Stevens S, Perera R, Holt TA, Salisbury C, et al. Factors associated with consultation rates in general practice in England, 2013–2014: a cross-sectional study. Br J Gen Pract. 2018; 68(670):e370–e7. Epub 2018/04/25. https://doi.org/10.3399/bjgp18X695981 PMID: 29686130; PubMed Central PMCID: PMC5916084.

12. Barnett K, Mercer SW, Norbury M, Watt G, Wyke S, Guthrie B. Epidemiology of multimorbidity and implications for health care, research, and medical education: a cross-sectional study. Lancet. 2012; 380(9836):37–43. Epub 2012/05/15. https://doi.org/10.1016/S0140-6736(12)60240-2 PMID: 22579043.

13. Pedersen KM, Andersen JS, Søndergaard J. General Practice and Primary Health Care in Denmark. The Journal of the American Board of Family Medicine. 2012; 25(Suppl 1):S34–S8. https://doi.org/10.3122/jabfm.2012.02.110216 %J The Journal of the American Board of Family Medicine. PMID: 22403249.

14. Kraft E, Hersperger M. Daten und Demographie–informativ und spannend. Schweizerische Ärztezeitung. 2009.

15. Lenoir AL, Richelle L, Ketterer F, Fraipont B, Cayn M, Duchesnes C, et al. Young general practitioners’ professional activities: a survey in the French-speaking part of Belgium. Acta clinica Belgica. 2017; 72(6):399–404. Epub 2017/03/21. https://doi.org/10.1080/17843286.2017.1302624 PMID: 28931747.

16. Hedden L, Barer ML, McGraill K, Law M, Bourgeault IL. In British Columbia, The Supply Of Primary Care Physicians Grew, But Their Rate Of Clinical Activity Declined. Health Aff (Millwood). 2017; 36(11):1904–11. Epub 2017/11/16. https://doi.org/10.1377/hlthaff.2017.0014 PMID: 29197511.

17. Hostetller S, Kraft E. FMH-Arzttestistik 2018. Schweizerische Ärztezeitung. 2019; 100(12):411–6.

18. Djalali S, Meier T, Hasler S, Rosemann T, Tandjung R. Primary care in Switzerland gains strength. Fam Pract. 2015; 32(3):348–53. https://doi.org/10.1093/fampra/cmv005 PMID: 25714346.

19. Federal Act on Research involving Human Beings 2014 [21.08.2019]. Available from: https://www.admin.ch/opc/en/classified-compilation/20061313/index.html.

20. Raumgliederungen: Federal Statistical Office; [cited 2019 23.07.2019]. Available from: https://www.bfs.admin.ch/bfs/de/home/grundlagen/raumgliederungen.html.

21. R Core Team. R: A language and environment for statistical computing. In: R Foundation for Statistical Computing, editor. Vienna, Austria2018.

22. Tandjung R, Hanhart A, Bartschi F, Keller R, Steinhauer A, Rosemann T, et al. Referral rates in Swiss primary care with a special emphasis on reasons for encounter. Swiss Med Wkly. 2015; 145:w14244. https://doi.org/10.4414/smw.2015.14244 PMID: 26709751.

23. Stevens S, Bankhead C, Mukhtar T, Perera-Salazar R, Holt TA, Salisbury C, et al. Patient-level and practice-level factors associated with consultation duration: a cross-sectional analysis of over one million consultations in English primary care. 2017; 7(11):e018261. https://doi.org/10.1136/bmjopen-2017-018261 PMID: 29150473.

24. Deveugele M, Derese A, van den Brink-Muinen A, Bensing J, De Maeseneer J. Consultation length in general practice: cross sectional study in six European countries. Bmj. 2002; 325(7362):472. Epub 2002/08/31. https://doi.org/10.1136/bmj.325.7362.472 PMID: 12203239; PubMed Central PMCID: PMC119444.

25. Fairchild DG, McLoughlin KS, Gharib S, Horsky J, Portnow M, Richter J, et al. Productivity, quality, and patient satisfaction: comparison of part-time and full-time primary care physicians. Journal of general internal medicine. 2001; 16(10):663–7. Epub 2001/10/27. https://doi.org/10.1111/j.1525-1497.2001.01111.x PMID: 11679033; PubMed Central PMCID: PMC1495282.

26. Bühren A, Eckert J. „Feminisierung“ der Ärztenschaft: Überschätzter Effekt. Dtsch Arztebl. 2011;108:A-968/B–968/C-968.

27. Hedden L, Barer ML, Cardiff K, McGraill KM, Law MR, Bourgeault IL. The implications of the feminization of the primary care physician workforce on service supply: a systematic review. Human resources for health. 2014; 12:32. Epub 2014/06/06. https://doi.org/10.1186/1478-4491-12-32 PMID: 24898264; PubMed Central PMCID: PMC4057816.

28. Pericin I, Mansfield G, Larkin J, Collins C. Future career intentions of recent GP graduates in Ireland: a trend analysis study. BJGP open. 2018; 2(1):bjgopen18X101409. Epub 2018/12/20. https://doi.org/10.3399/bjgopen18X101409 PMID: 30564707; PubMed Central PMCID: PMC6181082.
29. van Hassel D, van der Velden L, de Bakker D, Batenburg RJHRfH. Age-related differences in working hours among male and female GPs: an SMS-based time use study. 2017; 15(1):84. https://doi.org/10.1186/s12960-017-0258-4

30. McKinstry B, Colthart I, Elliott K, Hunter C. The feminization of the medical work force, implications for Scottish primary care: a survey of Scottish general practitioners. BMC Health Serv Res. 2006; 6:56. Epub 2006/05/12. https://doi.org/10.1186/1472-6963-6-56 PMID: 16686957; PubMed Central PMCID: PMC1475570.

31. Gisler LB, Bachofner M, Moser-Buchern CN, Scherz N, Streit S. From practice employee to (co-)owner: young GPs predict their future careers: a cross-sectional survey. BMC Family Practice. 2017; 18(1):12. Epub 2017/02/06. https://doi.org/10.1186/s12875-017-0591-7 PMID: 28148245; PubMed Central PMCID: PMC5289023.

32. van Hassel D, van der Velden L, de Bakker D, Batenburg R. Age-related differences in working hours among male and female GPs: an SMS-based time use study. Human resource for health. 2017; 15(1):84. https://doi.org/10.1186/s12960-017-0258-4 PMID: 29258573

33. Steinhaeuser J, Joos S, Szecsenyi J, Miksch A. A comparison of the workload of rural and urban primary care physicians in Germany: analysis of a questionnaire survey. BMC Family Practice. 2011; 12:112. Epub 2011/10/13. https://doi.org/10.1186/1471-2296-12-112 PMID: 21988900; PubMed Central PMCID: PMC3209467.

34. Canbek A. Es gibt keinen Arztmangel aufgrund «Feminisierung der Medizin». Schweiz Ärzteztg. 2019; 100(10):353–6 Epub 06.03.2019. https://doi.org/10.4414/saez.2019.17570.

35. Federal Statistical Office. (Projektübersicht) Gesundheitsversorgungsstatistik–Ambulante Statistiken im Projekt MARS. 2017.