Seasonal variations in the presentation and growth of thyroid cancer

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Summary Seasonal variation has been described in the presentation and growth of several malignant tumours, including cancers of the breast, uterus, uterine cervix, urinary bladder, liver, lymphatic system and skin, although the mechanisms are not known. We herein describe a circannual rhythm for thyroid cancer (total = 2627), with significantly more cases presenting during the late autumn and winter. In a subset of these cases (127 papillary carcinomas), significant seasonal variations with highest values in autumn–winter were found for tumour diameter and proliferation indicators (S- and G2,M-phase fractions). These results indicate the likelihood of a seasonal factor (or factors) of importance for the regulation and modification of tumour cell proliferation. When further clarified, this might be of relevance for the planning of diagnostic and therapeutic strategies.

Keywords: thyroid cancer; season; tumour diameter; proliferation; %S-phase–%G2,M-phase

Biological rhythms, especially circadian (24 h) and circannual (12 months), have been found in a wide range of physiological parameters and normal tissues (Aschoff, 1981; Shifrine et al, 1982; Halberg et al, 1983; Lærum et al, 1988; Sothern et al, 1995). Regarding seasonal variations, higher incidence of female breast cancer has been reported for the spring and summer (Cohen et al, 1983; Mason et al, 1985), and circannual contrasts have also been found for other tumours (Newell et al, 1985; Swerdlow, 1985; Hermida and Ayala, 1996). As for the thyroid, some studies have reported seasonal variations in endocrine parameters, with variables such as T₄, T₃ and thyroid-stimulating hormone (TSH) being higher in the autumn or winter (Halberg et al, 1981; Nicolau et al, 1987; Haus et al, 1988). As seasonal variations might be of potential interest for the detection and management of thyroid carcinomas, we wanted to review this group of tumours, most of which are slowly growing and have a good prognosis.

MATERIALS AND METHODS

Patients

During the period 1970–85, 2627 patients with thyroid cancer were reported to the Cancer Registry of Norway. Of these, 10% (n = 263) were surgically treated at the Department of Surgery, Haukeland University Hospital, in the period 1971–85. There were no major differences in the distribution of sex, age and histological types between our hospital cases and cases in the population-based Cancer Registry (Akslen and Myking, 1992). After histological re-examination of the 263 cases for exclusion of benign lesions and subtyping of malignant tumours according to the 1988 WHO criteria, 127 of these cases were found to represent papillary thyroid carcinomas with a diameter greater than 10 mm (microcarcinomas ≤ 10 mm were excluded). These 127 cases were included for further studies of seasonal variations with reference to time (month) of diagnosis, largest diameter of the primary tumour and proliferation indicators (S-phase and G2,M-phase fractions). Most of these patients (93%) were treated with total or near-total thyroidectomy. Regarding time of diagnosis, the date of initial histological/cytological diagnosis was used for both the registry (n = 2627) and hospital (n = 127) cases. The dates for the hospital subgroup are identical to those recorded in the Cancer Registry.

The mean age of the 91 female and 36 male patients was 46.0 years. Of the primary tumours (mean size ± s.e. = 29.7 ± 1.61 mm), 41% were completely intrathyroidal, 39% showed invasion into the thyroid gland capsule and 20% revealed major extrathyroidal extension, whereas 47% had lymph node metastases. Regarding histological grade (Akslen, 1993), 28% of these cases were classified as high-grade carcinomas. Distant metastases at diagnosis were present in two cases.

Data concerning regional and distant tumour recurrences and overall survival was recorded. Deaths from causes other than thyroid cancer were treated as censored observations. The median follow-up time was 137 months (maximum 271 months), and no patient was lost to follow-up.

Flow cytometric analysis and estimation of S-phase and G2,M-phase fractions were performed on paraffin-embedded tumour material as previously described (Akslen and Varhaug, 1995).

Data analysis

No differences in overall means for largest tumour diameter, %S-phase or %G2,M-phase were found between men and women by simple t-test, and data were subsequently combined for analyses. Monthly and seasonal tumour incidence were tested for time effect.
by chi-square \((n = 2627\) from national registry; \(n = 127\) local cases). Each data set consisting of monthly totals or means was analysed for circannual rhythm by the least-squares fit of a 1-year cosine using the single cosine procedure (Halberg et al., 1972), which has been adapted to the Macintosh computer (Mojón et al., 1992). In addition, harmonic components (6 months, 4 months) were added to the 1-year cosine model to see if a composite cosine would more accurately describe the true waveform of each data set (Portela et al., 1995), although at this time no biological significance can be ascribed to these additional components. The rhythm characteristics estimated by the cosinor procedure include the mesor (middle value of the fitted cosine representing the rhythm-adjusted mean), the amplitude (half the difference between the minimum and maximum of the fitted cosine) and the acrophase (time of peak value of fitted cosine). A \(P\)-value for rejection of the zero amplitude (no rhythm) assumption was determined, indicating whether or not the cosine model accounted for a significantly greater proportion of the variability in the time series when compared with the total variability around a flat line (the overall mean). Rhythm detection was considered statistically significant with a \(P\)-value of \(<0.05\). Although the cosinor method involving only a single fitted period may not accurately represent the true characteristics of the actual time-dependent variations if asymmetries exist in a time series, the procedure nevertheless is useful for objectively assessing and quantifying periodicities selected a priori – in our case, the year (Sothern, 1994). Because the mostly serially independent data were collected at unequidistant intervals over a number of years, the single cosinor method was thought to be the only procedure that could provide an objective estimate, not only of the circannual amplitude but also the circannual acrophase, i.e. time of peak value (Klemfuss and Clopton, 1993).

Associations were assessed by Pearson’s chi-square test. Univariate survival analysis (product-limit method) was performed using the BMDP-1L program, using the log-rank test for differences between groups. Recurrence-free survival, i.e. the time from diagnosis (in radically treated patients) until the first appearance of regional recurrences or distant spread, and overall patient survival (survival time until death from thyroid cancer) were compared across seasons.

**RESULTS**

**Registry cases**

The 2627 thyroid cancer cases diagnosed in Norway during 1970–85 showed a highly significant time-effect when incidence was compared across 12 months \((\chi^2 = 119.3, P<0.00001)\) or four seasons \((\chi^2 = 35.6, P<0.00001)\), being highest during the last 3-month period October–December (Figure 1A). A 1-year cosine analysis was not significant \((P = 0.163; Table 1)\), but a circannual pattern was prominent with highest values between September and January and fewer patients presenting during February–August, including the months of July and August when many Norwegians take vacation (Figure 1B). The addition of 6 or 4 months to the cosine model did not improve rhythm detection.

**Hospital cases**

The number of cases in our hospital series \((n = 127\) papillary carcinomas) showed a similar pattern of presentation, with 39, 32, 19 and 37 cases occurring in consecutive 3-month periods from January to December \((\chi^2\) for effect of season = 7.65, \(P = 0.05)\), with 60% of the cases occurring in the 6-month period between October and March. In these 127 cases, mean tumour diameter was 34.6 for men and 27.8 for women. This difference was not statistically significant \((P = 0.13)\). For all data combined, tumour diameter was greatest in the 3-month period of October–December \((mean \pm s.e. = 33.5\pm 3.6\, \text{mm})\), this value being 24% higher than in the 3-month period (July–September) with the lowest mean value \((27.0\pm 2.6\, \text{mm})\) (Figure 1C and D). Assuming a global form of the thyroid carcinomas, this corresponds to a mean increase in tumour volume of about 90% from lowest to highest average tumour size. There was a range of change (ROC) of 55% between lowest \((22.7\, \text{mm})\) and highest \((35.2\, \text{mm})\) monthly average tumour size (Table 1). A circannual rhythm in tumour size was detected at \(P = 0.040\) by single cosinor analysis of the 12 monthly means, with an acrophase (time of peak) in the late autumn (December; 95% limits: October 22 and January 25) and a double amplitude (representing a predictable range of change) of 30% (Table 1). The addition of 6 or 4 months to the cosine model did not improve rhythm detection.

There was no statistically significant difference between sexes for per cent of cells in S-phase \((women = 3.42 \pm 0.25\%\), men = \(3.69 \pm 0.71\%\); \(P = 0.29)\) or G,M-phase \((women = 3.94 \pm 0.24\%\), men = \(4.62 \pm 0.44\%\); \(P = 0.21)\). These proliferation indicators showed a similar circannual pattern, being highest during the last months of the year (Figure 1E–H). On average, mean S-phase fraction increased by 25% from a winter minimum of 3.04% \(\pm 0.43\%\) to an autumn maximum of 3.81% \(\pm 0.43\%\) and mean G,M-phase fraction increased by 30% from its spring minimum of 3.64% \(\pm 0.31\%\) to its fall maximum of 4.75% \(\pm 0.48\%\). Monthly means ranged from 2.43% \(\pm 0.36\%\) to 4.75% \(\pm 1.98\%\) for S-phase \((ROC = 95\%)\) and from 2.93% \(\pm 0.35\%\) to 5.21% \(\pm 0.80\%\) for G,M-phase \((ROC = 77\%)\). Cosinor analysis of monthly means showed a significant circannual rhythm for each proliferative indicator (Table 1; Figure 1E–H), with acrophases in the early autumn for S-phase \((P = 0.041,\ double\ amplitude = 31\%, \ acrophase = September 18)\) and late autumn for G,M-phase \((P = 0.012,\ double\ amplitude = 21\%, \ acrophase = November 3)\). For illustrative purposes, a composite cosine model consisting of 12+4 months, which seemed to better approximate secondary peaks and troughs for S-phase, and thus the observed waveform in these data, is also shown in Figure 1. The circannual component, however, remained the most prominent, and at this time no biological significance can be ascribed to the 4-month component. Of interest, a 4-month component has recently been used to better describe the circannual pattern for monthly uterine cervical cancer (Hermida and Ayala, 1996).

There was no significant seasonal variation for histological grade, primary tumour extension (pT stage) and presence of lymph node metastases (pN stage) when the 6-month period with maximum tumour diameter (October–March) was compared with the rest of the year. Further, analyses of successive 3-month periods showed no significant seasonal variations for these variables. The disease-free survival, as well as the overall survival of patients presenting during the period October–March (largest tumour diameter) was not different from those presenting during the rest of the year (log-rank test, \(P = 0.90\) and 0.40 respectively).
Figure 1  Seasonal variation (A, C, E and G) of thyroid cancer with respect to number of cases presenting in the Norwegian population (A and B, \( n = 2627 \)); tumour diameter (C and D), S-phase fraction (E and F) and G2M-phase fraction (G and H) in a subgroup (\( n = 127 \)) of the material (seasonal means \( \pm \) s.e.m.; some cases were excluded due to lack of information). The graphs in the right column show the corresponding monthly means with fitted 12-month cosine (and with fitted composite 12+4 month cosine when this model helped to describe the observed waveform of the data). Time scale = 3 monthly spans (A, C, E and G) or month (B, D, F and H). Highest values for each variable were found in autumn–winter, with lowest values in spring–summer.
Table 1: Circannual rhythm characteristics for thyroid cancer presentation and proliferation indicators observed in Norway, 1970–85

| Variable       | Units          | Total  | Monthly | Low   | High   | Absolute | Per cent | Mesor ± s.e. | P       | Amp(A) ± s.e. | (2A)* | Acrophase | (95% CL)          |
|----------------|----------------|--------|---------|-------|--------|----------|----------|-------------|--------|--------------|-------|-----------|------------------|
| Incidence      | Cases/months   | 2627   | 12      | 116   | 282    | 166      | 143%     | 222 ± 14    | 0.163  | 37 ± 17      | (33%) | December 19 | (–)               |
| Size           | Diameter (mm)  | 118    | 12      | 22.7  | 35.2   | 12.5     | 55%      | 28.8 ± 1.0  | 0.040  | 4.3 ± 1.4    | (30%) | December 8   | (October 22–January 25) |
| S-phase        | Per cent       | 120    | 12      | 2.43  | 4.75   | 2.32     | 95%      | 3.59 ± 0.14 | 0.041  | 0.56 ± 0.17  | (31%) | September 28 | (August 16–November 9) |
| G0M-phase      | Per cent       | 120    | 12      | 2.93  | 5.21   | 2.28     | 78%      | 4.04 ± 0.16 | 0.012  | 0.43 ± 0.10  | (21%) | November 3   | (October 1–December 6)  |

*Analysis of monthly totals or mean values. Single cosinor, least-squares fit of 1-year cosine to the 12 monthly values. Range of change, difference from lowest to highest value; Mesor, rhythm-adjusted 1-year mean; Amplitude (Amp), 1/2 peak-trough difference of fitted cosine (2A = double amplitude per cent of Mesor); Acrophase, peak of fitted cosine (in calendar month and day referenced from January 1); CL, confidence limits. P-value from non-zero amplitude test.

**DISCUSSION**

Our results indicate a significant circannual variation in the presentation of thyroid cancer. The distribution was highest in the late autumn and winter (September-December) and lowest in the early spring (March-May). A significant increase in the incidence of thyroid cancer occurs during these months, with a peak in December. This seasonal variation is consistent with previous studies reporting a higher incidence of thyroid cancer during the winter months (Halberg et al., 1983; Mason et al., 1985).

The highest incidence of thyroid cancer occurs in December, with a peak in 1990. This peak is followed by a slow decline in the early spring months. The incidence then rises again, peaking in the late autumn months (September-December). The highest incidence of thyroid cancer occurs in December, with a peak in 1990. This peak is followed by a slow decline in the early spring months. The incidence then rises again, peaking in the late autumn months (September-December).

Several factors should be considered to explain the seasonal variations in thyroid cancer presentation. These include the interaction of endogenous rhythms with external environmental factors, such as light exposure and temperature. The exact mechanisms by which these rhythms influence thyroid cancer presentation are not fully understood, but the results of our study suggest that the circannual rhythm may play a significant role in the seasonal variation of thyroid cancer presentation. Further research is needed to determine the specific mechanisms underlying this phenomenon.
women below 50 years of age (Cohen et al, 1983), suggesting a relationship to menstrual status.

Recent studies have indicated an independent relationship between time of presentation and the subsequent prognosis for patients with breast cancer, and those presenting during the spring/summer period might have the best prognosis (Mason et al, 1987; 1990). In our study of thyroid cancer, survival showed no significant variation with season of presentation.

Our findings support the view (Nicolau and Haus, 1992) that seasonal differences in DNA synthesis in normal and neoplastic tissues may be of clinical importance. Blank et al (1992) performed bone marrow biopsies on patients with various malignancies at different clock hours and calendar dates, and found that the mitotic activity revealed a significant circannual rhythm, with an amplitude of 37% and an acrophase occurring on 27 August, a timing nearly identical to that for DNA synthesis found in healthy bone marrows and rectal mucosa (Sotherrn et al, 1995). Preclinical studies have also shown a circannual tolerance effect for certain tissues after exposure to anti-cancer agents (Lévi et al, 1988; Mormont et al, 1988). Further, myelotoxicity was more severe in winter than in summer patients treated with chemotherapy for ovarian or bladder cancer (Hrushesky, 1982).

In conclusion, our present study adds further evidence to the existence of circannual rhythms in tumour pathology, and such results might be important for the timing of treatment regimens (for example dosing time of cytotoxic agents) and other procedures (for example the right season to examine/screen for early signs of malignancy). In addition, time of the year may also be relevant for epidemiological inference, especially when studying the influence of exogenous exposures on tumour development. Our findings on thyroid cancer seasonality suggest the presence of a seasonal factor (or factors) of importance to the regulation of tumour growth, although some influence of vacation patterns is also probable. When further clarified in larger studies and other locales, such data might be of relevance for the planning of diagnostic and therapeutic strategies.

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