BMJ Open

Systematic review and meta-analysis of the prognosis and prognostic factors of interstitial pneumonia with autoimmune features

Hiroyuki Kamiya1, Ogee Mer Panlaqui2

ABSTRACT

Objective To clarify the prognosis and prognostic factors of interstitial pneumonia with autoimmune features (IPAF) in comparison to idiopathic pulmonary fibrosis (IPF), the most common idiopathic interstitial pneumonia, and connective tissue disease-associated interstitial pneumonia (CTD-IP).

Design A systematic review and meta-analysis.

Data sources Electronic databases such as Medline and Embase were searched from 2015 through 6 September 2019.

Eligibility criteria for selecting studies Primary studies that comparatively investigated the prognosis or prognostic factors of IPAF were eligible.

Data extraction and analysis Two reviewers extracted relevant data and assessed the risk of bias independently. A meta-analysis was conducted using a random-effects model. The quality of presented evidence was assessed by the Grades of Recommendation, Assessment, Development, and Evaluation system.

Results Out of a total of 656 records retrieved, 12 studies were reviewed. The clinical features of IPAF were diverse between studies, which included a radiological and/or pathological usual interstitial pneumonia (UIP) pattern of between 0% and 73.8%. All studies contained some risk of bias. There was no significant difference of all-cause mortality between IPAF-UIP and IPF in all studies, although the prognosis of IPAF in contrast to IPF or CTD-IP varied between studies depending on the proportion of UIP pattern. Among the potential prognostic factors identified, age was significantly associated with all-cause mortality of IPAF by a pooled analysis of univariate results with a hazard ratio (HR) of 1.06 (95% confidence interval (CI) 1.04 to 1.07). The adjusted effect of age was also significant in all studies. The quality of presented evidence was deemed as very low.

Conclusion There was no significant difference of all-cause mortality between IPAF-UIP and IPF. Age was deemed as a prognostic factor for all-cause mortality of IPAF. The findings should be interpreted cautiously due to the low quality of the presented evidence.

PROSPERO registration number CRD42018115870.

INTRODUCTION

Interstitial pneumonia (IP) is a heterogeneous clinical entity that is pathologically characterised by varying degrees of a mixture of inflammation and fibrosis in the interstitium of pulmonary parenchyma. Although its pathogenesis is yet to be confirmed, IP is known to be accompanied by other medical conditions such as connective tissue disease (CTD), occupational dust exposure and drug toxicity. It is essential to identify these known causes and discriminate between idiopathic IPs (IIPs) and secondary IPs, because therapeutic strategies and prognosis for these diseases are different. Among these known aetiologies, CTDs are the most common underlying disease and are reported to be associated with approximately 30% of cases of IP that were initially diagnosed as IIPs. However, it is often difficult to make a definitive diagnosis of CTD-associated IP (CTD-IP) because IP can be a sole presenting manifestation of certain CTDs, or the symptoms and signs suggestive of CTDs may be too subtle to be recognised as an underlying disease. Furthermore, some cases remain undiagnosed as a defined CTD under current classification criteria regardless of the implication of autoimmune features. This latter case has been a subject of debate over the last decade because clinical differences between undifferentiated CTD-IP (UCTD-IP) and defined CTD-IP or IIP are closely related to

Strengths and limitations of this study

This study was a systematic review based on a comprehensive literature search to accumulate all of the evidence that is currently available.

The results were statistically synthesised to overcome the shortcoming of a small study, which is usually the case with reports of interstitial pneumonia with autoimmune features.

Most of the findings were inconclusive due to a small number of studies and the scarcity of multivariate analyses.
the management of patients with IP.10 However, different research groups have implemented different nomenclatures with diverse classification criteria to investigate this medical condition, and thus previous literature is flooded with inconsistent results derived from heterogeneous groups of patients who were recruited based on the same concept of the disease.11–14 In order to address these chaotic circumstances and facilitate a prospective study targeting a uniform cohort of patients who can be classified into a possibly new clinical entity, the European Respiratory Society (ERS) and the American Thoracic Society (ATS) Task Force proposed the new term ‘interstitial pneumonia with autoimmune features (IPAF)’ and reported its classification criteria in 2015.15 The criteria of IPAF are composed of clinical, serological and morphological domains (multiple features constituting each domain), and at least one feature from at least two of the domains each is required to be positive for the diagnosis of IPAF after confirming the presence of IP and excluding alternative aetiologies for the disease.15 Since the introduction of this terminology, many researchers have started using these classification criteria to describe patients with IP suggestive of autoimmune features who cannot be diagnosed as a defined CTD.16 Although the attempt of the Task Force was successful in this regard, the findings of recent reports remain diverse, and it is still unclear whether IPAF is a unique clinical entity from CTD-IP or IIP, in particular, idiopathic pulmonary fibrosis (IPF), which is the most common type of IIP.15 Therefore, the aim of this systematic review was to summarise previous reports of IPAF and present current evidence focusing on the prognosis and prognostic factors of IPAF in comparison to CTD-IP or IIPs.

METHODS
This review was conducted and reported according to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA)18 and the Meta-analysis of Observational Studies in Epidemiology (MOOSE) statement.19

Patient and public involvement
There was no patient or public involvement in the whole process of conducting this research.

Eligibility
This review included any type of primary study that compared the prognosis of IPAF with that of CTD-IP or IIPs such as IPF and idiopathic non-specific interstitial pneumonia, except for a case report. The studies that investigated the prognostic factors of IPAF for predefined outcomes were also eligible. IPAF was diagnosed based on the classification criteria proposed by the ERS/ATS Task Force in 2015, which consisted of CTD-specific features categorised into clinical, serological and morphological domains.15 IIPs were diagnosed based on the ATS/ERS statement,1 and a defined CTD was diagnosed based on its widely accepted classification criteria including the American College of Rheumatology criteria for rheumatoid arthritis.20–25 Although various terminologies had previously been implemented to describe patients who would have been classified as IPAF,11–14 they were excluded due to concerns about recruiting a potentially heterogeneous group of people regardless of their overlapping clinical features.26 There was no limitation on follow-up period, and patients were included at any point of time during their disease course. Potential prognostic factors were designated in this study and defined as any clinical information related to predefined outcomes such as demographic features, pulmonary functions, laboratory data and radiological and/or pathological patterns of IP.

Primary outcomes were all-cause and pulmonary-cause mortality. Secondary outcomes were a progression of the disease and deterioration of health-related quality of life. The progression of the disease was defined by each research group based on a combination of symptomatic, functional and/or radiological deterioration. The development of a defined CTD over the follow-up periods was also included as a secondary outcome.

Only English articles published in 2015 or later were eligible because the IPAF criteria were presented at the ERS congress and published in the European Respiratory Journal in that year.15 Editorials, letters and review articles were all ineligible, and conference proceedings were also excluded due to concerns about insufficient information.

Search of studies
Electronic databases, Medline (Ovid), Embase (Ovid) and Science Citation Index Expanded (Web of Science) were searched from 2015 through 24 December 2018. The search was updated on 6 September 2019. Because there had been no established subject heading for IPAF, other relevant terms such as ‘interstitial lung disease’ and ‘undifferentiated connective tissue disease’ were used and combined with text words of the study population to construct a string of search terms (eg, (‘interstitial lung disease’ or ‘interstitial pneumonia’ or ‘pulmonary fibrosis’) AND (‘undifferentiated connective tissue disease’ or ‘interstitial pneumonia with autoimmune features’)). Methodology filters such as ‘prognosis’ were not used to increase the sensitivity of the search (online supplementary e-Appendix). Grey literature was sought through Google Scholar.27 Reference lists of eligible studies and relevant review articles were hand searched, and cross referencing was also used to identify additional reports.

Study selection and data collection
Two reviewers (HK and OMP) independently examined the titles and abstracts of all retrieved articles after removing duplicates and selected eligible studies. Data were also extracted by the same reviewers based on a data extraction form, which was adopted and modified from a previously published article.28 Any disagreement in these
processes was resolved through discussion between the reviewers. Extracted data included the first author name, publication year, study location, study design, the number of participants and their demographic features, comparative groups, prognostic factors, outcomes, methods for statistical analysis, summary statistics and items associated with a risk of bias. Authors were asked to provide missing data for relevant outcomes.

### Risk of bias in individual studies

Risk of bias in individual studies was assessed by the Quality in Prognostic Studies tool. It consisted of six domains, which were rated as either low, moderate or high, and a study with low risk of bias in all domains was designated as a study with low risk of bias.

### Statistical analysis

#### Summary statistics

If prognosis of IPAF was compared with that of CTD-IP or IIPs by the Kaplan-Meier survival curve and the log rank test, the result was to be converted into a hazard ratio (HR) if possible, or otherwise the original data were presented. If the outcome was binary, the effect size of potential prognostic factors of IPAF was to be summarised as either an odds ratio (OR), risk ratio (RR) or HR. If the outcome or potential prognostic factors were continuous, they were to be summarised as the mean difference or the difference of the median value between two comparative groups, although there were no such cases in this review. The analysis focused only on a comparison of the prognosis and prognostic factors of IPAF reported in at least three studies given the assumption that frequent reports would indicate clinical relevance.

#### Meta-analysis

The results were combined only if the effect of a certain potential prognostic factor had been reported for the same outcome with the same summary statistics in at least three studies. Only univariate results were combined if the meta-analysis was appropriate because multivariate results were derived from different models constructed with multiple variables, which were diverse between studies. Meta-analysis was performed by a random-effects model with the DerSimonian and Laird method using statistical software Review Manager (RevMan) V.5.3 (Copenhagen: The Nordic Cochrane Centre, The Cochrane Collaboration, 2014). All of the results were presented with the 95% confidence interval (CI) in conjunction with the 95% prediction interval (PI), which was calculated if a meta-analysis was conducted and there was heterogeneity between studies. Statistical significance was set at p<0.05.

#### Heterogeneity and additional analysis

Heterogeneity was assessed by the Q statistic and I square. Statistical significance of heterogeneity was set at p<0.1, and the magnitude was interpreted as not important (0%–30%), moderate (30%–50%), substantial (50%–70%) or considerable (70%–100%). IPAF was likely to include a heterogeneous group of patients because it was diagnosed based on a combination of positive features in multiple domains. A different combination of positivity may have yielded a group with different clinical characteristics. Furthermore, a different number of cases with a usual interstitial pneumonia (UIP) pattern in IPAF was thought to affect the clinical course of the disease because UIP was noted to demonstrate a poor prognosis compared with other radiological and/or pathological patterns of IIPs. Therefore, the results were to be reanalysed depending on the proportion of radiological and/or pathological UIP pattern if data were available. Another subgroup analysis was also considered based on studies with a similar combination of positive domains, which contributed to the diagnosis of IPAF. Sensitivity analysis was to be conducted focusing on studies with low risk of bias alone. Small study bias including publication bias was to be examined graphically by a funnel plot and statistically by the Egger’s test if 10 or more studies were available for meta-analysis. The test was to be performed using Stata V.14 (STATA Corp LLC) with statistical significance of p<0.1.

#### Confirmation of prognostic factors

Prognostic factors were confirmed based on the results of multivariate analyses. If the effect of a potential prognostic factor demonstrated the same trend in all studies and was significant in the majority of multiple studies (>50%), it was deemed as a prognostic factor.

#### Assessment of quality of presented evidence

The Grades of Recommendation, Assessment, Development, and Evaluation (GRADE) system was applied to evaluate the quality of evidence regarding identified prognostic factors of IPAF.

### RESULTS

#### Search and selection of eligible reports

After removing 181 duplicates from a total of 656 records retrieved through a search of 4 electronic databases, 475 records were screened by the titles and abstracts, which identified 109 ineligible articles composed of case reports (n=8), conference abstracts (n=58), letters or editorials (n=27), non-English reports (n=5) and review articles (n=11). After 335 more records were excluded due to irrelevant subjects, the remaining 31 records were retrieved as full texts, and finally 12 reports/studies were eligible for this study (figure 1). Among them, three studies were conducted by the same research group. No additional study was identified by searching the reference lists or cross referencing of eligible reports and relevant review articles.

#### Characteristics of included studies

All studies were of retrospective cohort design (table 1). In seven studies, all types of IPs were screened to identify cases of IPAF (Chung et al, Oldham et al, Adegunsoye et al, Ahmad et al, Collins et al, Dai et al, Lim et al), whereas in two studies, IPAF was selected from cases of IP
Figure 1  Study flow diagram A total of 656 records were retrieved through a search of four electronic databases, Medline, EMBASE, Science Citation Index Expanded, and Google Scholar. After removing 181 duplicates, 475 records were screened by titles and abstracts, which identified 109 ineligible articles composed of case reports (n=8), conference abstracts (n=58), letters or editorials (n=27), non-English articles (n=5) and traditional review papers (n=11). After 335 more records were excluded due to irrelevant subjects, the remaining 31 records were retrieved as full texts, and finally 12 reports/studies, which included 3 studies conducted by the same research group, were eligible for this study. IPAF, interstitial pneumonia with autoimmune features.
Table 1  Characteristics of included studies

| Study            | Location        | Design       | Number | Sex (male/female %) | Age (years) | Ever smoker | Follow-up (years) | Clinical domain | Serological domain | Morphological domain | UIP pattern* | Outcome                  | Comparison†       | Statistical method |
|------------------|-----------------|--------------|--------|---------------------|-------------|-------------|-------------------|------------------|-------------------|---------------------|---------------|--------------------------|-------------------|---------------------|
| Chung et al. 2016 | USA  Retrospective cohort | 136         |        | (49.3/50.7)       | 63.5±10.9   | –           | –                 | –                | –                 | –                   | –             | 70/136 (51.5)  All-cause mortality | IPAF             | Cox                 |
| Oldham et al. 2016 | USA  Retrospective cohort | 144         |        | 69/75 (47.9/52.1) | 63.2±11.0   | 79/144(54.5) | –                 | 71/144 (49.3)    | 132/144 (91.7)   | 123/144 (85.4) | 77/141 (54.6)  All-cause mortality | IPAF             | Cox                 |
| Adegunsoye et al. 2017 | USA  Retrospective cohort | 84          |        | 41/43 (48.8/51.2) | 60.1±10.0   | 47/84 (56.0) | –                 | 36/- (44.4)     | 79/84 (94.0)     | –                   | –             | 62/84 (73.8)  All-cause mortality | IPAF             | Cox                 |
| Ahmad et al. 2017 | France  Retrospective cohort | 57          |        | 29/28 (50.9/49.1) | 64.4±14.0   | 16/47 (34.0) | 1.5 (median)        | 27/57 (47.3)    | 53/57 (93.0)     | 45/57 (78.9)  | 16/54 (29.6)  All-cause mortality | IPAF             | Cox                 |
| Collins et al. 2017 | USA  Retrospective cohort | 15          |        | 8/7 (53.3/48.7)   | 54.6±11.8   | 7/15 (46.7)  | –                 | –                | –                 | –                   | –             | 5/15 (33.3)  Change of PFT   | IPAF versus IF | Tukey               |
| Dai et al. 2018   | China  Retrospective cohort | 177         |        | 78/99 (44.1/55.9) | 60.2±12.9   | 34/177 (19.2) | –                 | 36/177 (20.3)   | 163/177 (92.1)   | 169/177 (95.9) | 8/177 (4.5)  All-cause mortality | IPAF             | Cox                 |
| Lim et al. 2019   | Korea  Retrospective cohort | 54          |        | 19/35 (35.2/64.8) | 67.9±10.5   | 15/54 (27.8) | –                 | 17/54 (31.9)    | 40/54 (80.7)    | 44/54 (81.5)  | 15/54 (27.8)  All-cause mortality | IPAF versus IP | Log rank           |
| Ito et al. 2017   | Japan  Retrospective cohort | 98          |        | 41/57 (41.8/58.2) | 67.5 (59.0–76.0) | 38/98 (38.8) | 4.6 (median)       | –                | 98/98 (100)     | 98/98 (100) | 98/98 (100) | 0/98 (0)          | All-cause mortality | IPAF             | Cox                 |
| Mejia et al. 2017 | Mexico  Retrospective cohort | 34          |        | 8/26 (23.5/76.5)  | 53.2±9.3    | –           | –                 | –                | –                 | –                   | –             | 1/34 (2.9)   All-cause mortality | IPAF versus CTD-IP† | Log rank           |
| Kelly and Moua 2017 | USA  Retrospective cohort | 101         |        | 40/61 (39.6/60.4) | 56.9±14.3   | 31/101 (30.7) | –                 | 77/101 (76.2)   | 87/101 (86.1)   | 96/101 (95.0) | 19/101 (18.8) All-cause mortality | IPAF             | Log rank           |

doi:10.1136/bmjopen-2019-031444
Continued
### Table 1

| Study              | Location | Design          | Number | Sex (male/female (%)) | Age (years) | Ever smoker (years) | Follow-up (years) | Clinical domain | Serological domain | Morphological domain | UIP pattern* | Outcome | Comparison† | Statistical method |
|--------------------|----------|-----------------|--------|-----------------------|-------------|---------------------|------------------|------------------|-------------------|---------------------|--------------|---------|-------------|-------------------|
| Yoshimura et al. 2018¹ ⁷ | Japan    | Retrospective cohort | 32     | 19/13 (59.4/40.6)     | 63.4±12.6   | 18/32 (56.2)        | 5.6 (mean)       | 17/32 (53.1)     | 23/32 (71.9)     | 31/32 (96.9)       | 13/32 (40.6) | All-cause mortality | IPAF versus CFIP  
|                    |          |                 |        |                       |             |                     |                  |                  |                   |                     |            |          |            | Cox               |
| Kim et al. 2019¹ ⁸ | Korea    | Retrospective cohort | 109    | 48/61 (44.0/56.0)     | 60.6±11.5   | 40/109 (36.7)       | –                | 49/109 (45.0)    | 88/109 (80.7)    | 100/109 (91.7)    | 40/109 (36.7) | All-cause mortality | IPAF versus CTD-IP††  
|                    |          |                 |        |                       |             |                     |                  |                  |                   |                     |            |          |            | Log rank          |
|                    |          |                 |        |                       |             |                     |                  |                  |                   |                     |            |          |            |                   |
|                    |          |                 |        |                       |             |                     |                  |                  |                   |                     |            |          |            |                   |
|                    |          |                 |        |                       |             |                     |                  |                  |                   |                     |            |          |            |                   |

Data are expressed as means±SD or counts/total number (percent) unless otherwise specified.

*All studies were based on the finding of HRCT scan excluding two studies (Kelly and Moua,
¹⁵ Kim et al.,¹⁸) which additionally considered pathological findings and another study (Adegunsoye et al.,³⁹) which was based on pathological findings alone.
†‘IPAF’ indicating a comparative analysis between IPAF with and without potential prognostic factors.
‡Median (range).§No detailed explanation for CTD-IP or CTD-UIP.
¶Scleroderma spectrum disease and RA were the most common, accounting for 37.1% and 22.9%, respectively and a radiological UIP pattern on HRCT was present in 33% of the total number of patients.
**Underlying CTDs were comprised of 46 cases of RA, 18 of SSc, 6 of Addison syndrome, 5 of inflammatory muscle disease and 7 of SLE.
††314 cases of underlying CTDs were idiopathic inflammatory myopathy spectrum disease and no UIP pattern was confirmed on HRCT.
†‡Underlying CTDs were comprised of 66 cases of RA, 25 of PM/DM, 24 of SSc, 7 of MCTD, 6 of overlap syndromes and 2 of SLE.
¶¶IPAF comprised of 333 cases of IPF, 36 cases of NSIP, 35 cases of COP and 73 cases of unclassifiable IP.
§§IPAF-NSIP comprised of 333 cases of IPF, 36 cases of NSIP, 35 cases of COP and 73 cases of unclassifiable IP.
|                    |          |                 |        |                       |             |                     |                  |                  |                   |                     |            |          |            |                   |

Change of PFT

CTD incidence

IPAF

Descriptive

IPAF versus CFIP  
IPAF-UIP versus IPF  
IPAF-NSIP versus iNSIP  
IPAF versus CTD-IP††  
IPAF versus COP††  
IPAF versus NSIP§  
IPAF versus UIP§  
IPAF versus IIP§  
IPAF versus nonUIP§  
IPAF-UIP versus IIP§  
CFIP, chronic fibrosing interstitial pneumonia; COP, cryptogenic organizing pneumonia; CTD-IP, connective tissue disease-associated interstitial pneumonia; HRCT, high-resolution computed tomography; IP, idiopathic interstitial pneumonia; NSIP, idiopathic non-specific interstitial pneumonia; UIP, usual interstitial pneumonia; IPAF, interstitial pneumonia with autoimmune features; IPF, idiopathic pulmonary fibrosis; MCTD, mixed connective tissue disease; PFT, pulmonary function test; PM/DM, polymyositis/dermatomyositis; RA, rheumatoid arthritis; SLE, systemic lupus erythematosus; SSc, systemic sclerosis; UIP, usual interstitial pneumonia.
with positive serology of IPAF classification criteria and myositis-specific autoantibodies, respectively (Ito et al, Mejía et al). The other three studies identified cases of IPAF from a cohort of UCTD-IP, chronic fibrosing interstitial pneumonia (CFIP) and IIPs, respectively (Kelly and Moua, Yoshimura et al, Kim et al).

The mean or median follow-up period was reported in only three studies, which spanned over 1.5–5.6 years. The number of participants ranged from 15 to 177, and their mean or median ages were between 53.2 and 67.9 years. The proportion of men and ever smokers was distributed between 23.5% and 59.4% and 19.2% and 56.2%, respectively. Positivity of clinical, serological and morphological domains of IPAF classification criteria ranged from 20.3% to 76.2%, 71.9% to 100% and 78.9% to 100%, respectively. UIP pattern was confirmed on a high-resolution computed tomography (HRCT) scan and/or a surgical lung biopsy (SLB) for between 0% and 73.8% of cases of IPAF. All-cause mortality was the primary outcome in all studies except for one (table 1).

### Risk-of-bias assessment

All studies included in the review were deemed as containing some risk of bias. In particular, study attrition and statistical analysis were designated as a high risk of bias. The attrition of participants over time was unclear in all studies. Statistical analysis and reporting were also insufficient due to the lack of explanation of the analytical strategy and model development (table 2).

### Prognosis of IPAF in comparison to other IPs

All-cause mortality of IPAF was significantly better than that of IPF in two studies (Dai et al, Lim et al) (19.8% vs 31.9%, p<0.001 and 27.8% vs 63.4%, p=0.008, respectively), which included UIP pattern on HRCT in 4.5% and 27.8% of the cases, respectively (table 3). However, there was no significant difference of all-cause mortality in the other two studies (Oldham et al, Ahmad et al), which included a larger number of cases with UIP pattern on HRCT (54.6% and 29.6%, respectively). There was no significant difference of all-cause mortality between IPAF-UIP and IPF in all five studies that reported it (Oldham et al, Lim et al, Kelly and Moua, Yoshimura et al, Kim et al; table 3). Among the four studies comparing IPAF with CTD-IP, only one study (Oldham et al) demonstrated higher all-cause mortality for IPAF (p<0.001). It included a larger proportion of a radiological and/or pathological UIP pattern (54.6%) in contrast to 27.8%, 2.9% and 36.7% in the other three studies (Lim et al, Mejía et al, Kim et al), respectively.

### Univariate analysis of potential prognostic factors

Age, men, ever smokers, radiological and/or pathological UIP pattern, percentage of predicted forced vital capacity (%FVC), and percentage of predicted diffusing capacity of the lung for carbon monoxide (%DLCO) were identified as potential prognostic factors for all-cause mortality of IPAF, which were reported in at least three studies (table 4).

Age was significantly associated with all-cause mortality of IPAF in all four studies that reported it (Oldham et al, Dai et al, Ito et al, Kim et al; table 4), which produced a combined effect of an HR of 1.06 (95% CI 1.04 to 1.07) with no heterogeneity (figure 2). There was also a trend of worse all-cause mortality for men than women in all of these studies, which generated a significant result by a pooled analysis with an HR of 1.66 (95% CI 1.22 to 2.25)

## Table 2 Risk of bias in included studies

| Study               | Study participation | Study attrition | Prognostic factor measurement | Outcome measurement | Study confounding | Statistical analysis and reporting |
|---------------------|---------------------|-----------------|-------------------------------|---------------------|-------------------|----------------------------------|
| Chung et al 2016    | Low risk            | High risk       | Low risk                      | Low risk            | Moderate risk      | High risk                       |
| Oldham et al 2016   | Low risk            | High risk       | Low risk                      | Low risk            | Moderate risk      | High risk                       |
| Adegunsoye et al    | Moderate risk       | High risk       | Low risk                      | Low risk            | Moderate risk      | High risk                       |
| Ahmad et al 2017    | Low risk            | High risk       | Low risk                      | Low risk            | Moderate risk      | High risk                       |
| Collins et al 2017  | Low risk            | High risk       | Low risk                      | Low risk            | High risk         | High risk                       |
| Dai et al 2018      | Low risk            | High risk       | Low risk                      | Low risk            | High risk         | High risk                       |
| Lim et al 2019      | Low risk            | High risk       | Low risk                      | Low risk            | Moderate risk      | High risk                       |
| Ito et al 2017      | Low risk            | High risk       | Low risk                      | Low risk            | Moderate risk      | High risk                       |
| Mejia et al 2017    | Low risk            | High risk       | Low risk                      | Low risk            | Moderate risk      | High risk                       |
| Kelly and Moua 2017 | Low risk            | High risk       | Low risk                      | Low risk            | Low risk          | High risk                       |
| Yoshimura et al    | Low risk            | High risk       | Low risk                      | Low risk            | Moderate risk      | High risk                       |
| Kim et al 2019      | Low risk            | High risk       | Low risk                      | Low risk            | Moderate risk      | High risk                       |

Bold text indicates high risk of bias.
with no heterogeneity (figure 3). The effect of a smoking habit was reported in five studies (Oldham et al, Ahmad et al, Dai et al, Ito et al, Kim et al), and a pooled analysis of these studies demonstrated that ever smokers were not significantly associated with all-cause mortality with an HR of 1.54 (95% CI 0.97 to 2.46) and substantial heterogeneity (figure 4). One study (Oldham et al) demonstrated the opposite effect to the others and seemed to be responsible for some of the heterogeneity. A UIP pattern was noted in the majority of IPAF (54.6%) in that study in contrast to the other four studies, which demonstrated the proportion of UIP pattern as 29.6%, 4.5%, 0% and 36.7%, respectively. The effect of a radiological and/or pathological UIP pattern was reported in six studies (table 4). Because Oldham et al and Adegunsoye et al were conducted by the same research group sharing the same cohort, the former study was prioritised for the analysis considering its larger sample size. A pooled analysis of three out of these five studies (Oldham et al, Dai et al, Kim et al) demonstrated that a radiological and/or pathological UIP pattern was significantly associated with all-cause mortality of IPAF with an HR of 2.93 (95% CI 1.65 to 5.20) and moderate heterogeneity (figure 5). Only one study (Dai et al) demonstrated a non-significant result, and it included the least proportion (4.5%) of IPAF subjects with a radiological and/or pathological UIP pattern in contrast to 54.6% and 36.7% in the other two studies (Oldham et al and Kim et al), respectively. A pooled analysis of three studies (Oldham et al, Ito et al and Kim et al) demonstrated that both %FVC and %DLCO were significantly associated with all-cause mortality of IPAF with HRs of 0.89 (95% CI 0.81 to 0.98) and 0.77 (95% CI 0.69 to 0.85), respectively, with no heterogeneity (figures 6 and 7).

**Multivariate analysis of potential prognostic factors**

The adjusted effect of age on all-cause mortality of IPAF was reported in four studies, all of which demonstrated significant results and almost the same effect size with HRs ranging from 1.04 to 1.07 (table 4). Men were not significantly associated with all-cause mortality with an HR of 1.65 (95% CI 0.92 to 2.97) in one study (Oldham et al). The adjusted effect of smoking habit on all-cause mortality was reported in two studies. One of these studies showed a significant result with an HR of 2.11 (95% CI 1.02 to 4.36; Dai et al), whereas the other study showed a non-significant result with an HR of 1.11 (95% CI 0.60 to 2.05; Oldham et al). The effect of a radiological and/or pathological UIP pattern was reported using multivariate analyses in three studies (Oldham et al, Adegunsoye et al and Kim et al). Oldham et al was prioritised over Adegunsoye et al for the analysis considering its larger sample size in the same cohort, which demonstrated a non-significant result with an HR of 1.72 (95% CI 0.83 to 3.56). However, another study (Kim et al) reported a significantly worse all-cause mortality with a radiological and/or pathological UIP pattern with an HR of 3.85 (95% CI 1.99 to 7.43). The effect of %FVC and %DLCO was reported using a multivariate analysis in only one study each. There was no significant association between %FVC and all-cause mortality with an HR of 1.00 (95% CI 0.78 to 1.29; Oldham et al), whereas %DLCO was significantly associated with all-cause mortality with an HR of 0.66 (95% CI 0.57 to 0.78; Kim et al).

---

**Table 3** All-cause mortality of interstitial pneumonia with autoimmune features (IPAF) compared with other interstitial pneumonias

| Study | Effect size† |
|-------|--------------|
| IPAF versus IPF | Oldham et al 2016 | P=0.07 (log rank) |
| | Ahmad et al 2017 | 16.4% versus 5.2%, p=0.05 (log rank) |
| | Dai et al 2018 | 19.8% versus 31.9%, p<0.001 (log rank) |
| | Lim et al 2019 | 27.8% versus 63.4%, p=0.008 (Cox) |
| IPAF-UIP‡ versus IPF | Oldham et al 2016 | P=0.51 (log rank) |
| | Lim et al 2019 | Mean ST 64.6 versus 52.0 months, p=0.08 (log-rank) |
| | Kelly and Moua 2017 | P=0.08 (log rank) |
| | Yoshimura et al 2018 | P=0.09 (log rank) |
| | Kim et al 2019 | Median ST 36 versus 51 months, p=0.43 (log rank) |
| IPAF versus CTD-IP | Oldham et al 2016 | P=0.001 (log rank)§ |
| | Lim et al 2019 | 27.8% versus 21.1%, p=0.41 (Fisher) |
| | Mejía et al 2017 | P>0.83 (log rank) |
| | Kim et al 2019 | Median ST 160 versus 142 months, p=0.98 (log rank) |

†Only a comparison of all-cause mortality of IPAF with that of other interstitial pneumonias, which was reported in at least three studies, was summarised. Italic bold indicates statistical significance.

‡Only p value was described if effect estimates were not available.

§IPAF pattern was diagnosed based on radiological and/or pathological findings in two studies (Kelly and Moua, Kim et al) whereas it was determined radiologically in the other studies (Oldman et al, Lim et al, Yoshimura et al).

§IPAF demonstrated a higher mortality than CTD-IP.

CTD-IP, connective tissue disease-associated interstitial pneumonia; IPF, idiopathic pulmonary fibrosis; ST, survival time; UIP, usual interstitial pneumonia.
Table 4  Potential prognostic factors for interstitial pneumonia with autoimmune features (IPAF)*

| Potential prognostic factors† | Univariate | | | Multivariate | | |
| --- | --- | --- | --- | --- | --- | --- |
| | Study | Effect size (HR (95% CI)) | Study | Effect size (HR (95% CI)) | | |
| **Age (/1 year)** | | | | | | |
| | Oldham et al 201638 | 1.06 (1.03 to 1.08) | Oldham et al 201638 | 1.05 (1.02 to 1.08) | | |
| | Dai et al 201840 | 1.04 (1.01 to 1.07) | Dai et al 201842 | 1.07 (1.02 to 1.12) | | |
| | Ito et al 201744 | 1.07 (1.03 to 1.11) | Ito et al 201744 | 1.04 (1.01 to 1.06) | | |
| | Kim et al 201948 | 1.06 (1.03 to 1.09) | Kim et al 201948 | 1.06 (1.03 to 1.10) | | |
| **Men (vs women)** | | | | | | |
| | Oldham et al 201638 | 1.54 (0.91 to 2.60) | Oldham et al 201638 | 1.65 (0.92 to 2.97) | | |
| | Dai et al 201842 | 1.89 (0.94 to 3.80) | | | | |
| | Ito et al 201744 | 1.98 (0.92 to 4.25) | | | | |
| | Kim et al 201948 | 1.50 (0.87 to 2.60) | | | | |
| **Ever smoker (vs never smoker)** | | | | | | |
| | Oldham et al 201638 | 0.92 (0.54 to 1.54) | Oldham et al 201638 | 1.11 (0.60 to 2.05) | | |
| | Ahmad et al 201740 | 7.18 (1.31 to 39.26) | | | | |
| | Ito et al 201744 | 2.44 (1.19 to 5.00) | | | | |
| | Kim et al 201948 | 1.27 (0.59 to 2.73) | | | | |
| | Oldham et al 201638 | 1.50 (0.87 to 2.60) | | | | |
| **UIP (HRCT and/or SLB) (vs nonUIP)**‡ | | | | | | |
| | Oldham et al 201638 | 2.40 (1.21 to 4.76) | Oldham et al 201638 | 1.72 (0.83 to 3.56) | | |
| | Adegunsoye et al 201739 | 6.11 (1.43 to 26.08) | Adegunsoye et al 201739 | 5.35 (11.21 to 23.67) | | |
| | Ahmad et al 201740 | p=0.23 (log-rank)§ | Kim et al 201948 | 3.85 (1.99 to 7.43) | | |
| | Dai et al 201842 | 1.55 (0.47 to 5.11) | | | | |
| | Kelly and Moua 201738 | p=0.04 (log-rank)¶ | | | | |
| | Kim et al 201948 | 4.47 (2.54 to 7.88) | | | | |
| **Percentage of predicted FVC (/10%)** | | | | | | |
| | Oldham et al 201638 | 0.90 (0.78 to 1.05) | Oldham et al 201638 | 1.00 (0.78 to 1.29) | | |
| | Ito et al 201744 | 0.82 (0.67 to 1.00) | | | | |
| | Kim et al 201948 | 0.93 (0.80 to 1.09) | | | | |
| **Percentage of predicted DLCO (/10%)** | | | | | | |
| | Oldham et al 201638 | 0.74 (0.63 to 0.86) | Kim et al 201948 | 0.66 (0.57 to 0.78) | | |
| | Ito et al 201744 | 0.90 (0.70 to 1.16) | | | | |
| | Kim et al 201948 | 0.75 (0.64 to 0.87) | | | | |

*Italic bold indicates statistical significance.
†Potential prognostic factors were defined as any clinical information that was reported by at least three studies.
‡UIP pattern was confirmed based on the finding of HRCT scan in three studies (Oldham et al,38 Ahmad et al,40 Dai et al,42), pathological finding in one study (Adegunsoye et al39) and either findings in two studies (Kelly and Moua,48 Kim et al444).
§Survival was estimated by the Kaplan-Meier survival curve and the log-rank test demonstrated that the prognosis of IPAF–UIP tended to be worse than that of IPAF–nonUIP.
¶Survival was estimated by the Kaplan-Meier survival curve and the log-rank test demonstrated that the prognosis of IPAF–UIP was significantly worse than that of IPAF–nonUIP.

Figure 2  Forrest plot of the effect of age on all-cause mortality of interstitial pneumonia with autoimmune features. The effect of age on all-cause mortality was reported in four studies, in which a total of 528 participants with 236 men (44.7%) were included. Meta-analysis of these studies demonstrated that age was significantly associated with all-cause mortality with an HR of 1.06 (95% CI 1.04 to 1.07, p<0.00001) with no heterogeneity.
Figure 3  Forrest plot of the effect of sex on all-cause mortality of interstitial pneumonia with autoimmune features. The effect of sex on all-cause mortality was reported in four studies, in which a total of 528 participants with 236 men (44.7%) were included. Meta-analysis of these studies demonstrated that all-cause mortality of men was significantly worse than that of women with an HR of 1.66 (95% CI 1.22 to 2.25, p=0.001) with no heterogeneity.

Additional analysis
Neither a subgroup analysis nor sensitivity analysis was conducted due to the small number of studies and some risk of bias identified in all studies. Small study bias could not be examined due to the small number of studies.

Confirmation of prognostic factors
Among six potential prognostic factors identified, only age was deemed to contribute to all-cause mortality of IPAF. It was significant by multivariate analysis in all studies, and all of the results showed a similar trend.

Quality of evidence
Quality of evidence for age as a prognostic factor for IPAF was rated as very low by the GRADE system (table 5).

DISCUSSION
This systematic review and meta-analysis demonstrated that the clinical features of IPAF, such as demographics, number of cases with a UIP pattern and positivity of each domain constituting the classification criteria of IPAF, were diverse between studies. All-cause mortality of IPAF/UIP was not significantly different from that of IPF in all studies, although the prognosis of IPAF in comparison to IPF or CTD-IP was inconsistent between studies and seemed to have been related to the proportion of cases with a radiological and/or pathological UIP pattern. Only age was deemed as a prognostic factor for all-cause mortality of IPAF based on multivariate results, although sex, smoking habit, radiological and/or pathological UIP pattern, %FVC and %DLCO were also considered as potential prognostic factors.

The variability of clinical features of IPAF between studies was described in a previous review. It may be mostly related to the retrospective design of studies and a different process of identifying cases of IPAF. In seven studies, all types of IP were screened to identify cases of IPAF, whereas two studies sought to identify IPAF from cases of IP with positive serology of IPAF classification criteria and myositis-specific autoantibodies, respectively. The other three studies identified cases of IPAF from UCTD-IP, CFIP and IIPs, respectively. As a result, this potential selection bias may have caused clinical heterogeneity between the studies in this review. The presence of a radiological and/or pathological UIP pattern is particularly important for the analysis of prognosis of IPAF because it is reported to be predictive of a worse prognosis of IIPs. It is still unclear whether IPAF represents

Figure 4  Forrest plot of the effect of smoking habit on all-cause mortality of interstitial pneumonia with autoimmune features. The effect of smoking habit was reported in five studies, in which a total of 585 participants with 265 men (45.3%) were included. Meta-analysis of these studies demonstrated that ever smokers were not significantly associated with all-cause mortality with an HR of 1.54 (95% CI 0.97 to 2.46, p=0.07/95% prediction interval 0.37 to 6.33). There was substantial heterogeneity between the results of the included studies ($\chi^2=8.26$, df=4, p=0.08, $I^2=52\%$). Only one study (Oldham et al) demonstrated an opposite effect to the others, and the majority of participants were ever smokers with a radiological usual interstitial pneumonia pattern in that study.
Figure 5  Forrest plot of the effect of a usual interstitial pneumonia (UIP) pattern on a high-resolution computed tomography (HRCT) and/or a surgical lung biopsy (SLB) specimen on all-cause mortality of interstitial pneumonia with autoimmune features (IPAF). The effect of UIP pattern on all-cause mortality was reported in six studies. Meta-analysis was conducted for three of these studies, in which a total of 430 participants with 195 men (45.3%) were included. A UIP pattern on HRCT and/or SLB was significantly associated with all-cause mortality of IPAF with an HR of 2.93 (95% CI 1.65 to 5.20, \( p=0.0002/95% \) prediction interval 0.01 to 808.7). There was moderate heterogeneity between the results of the included studies (\( \chi^2=3.44, \text{df}=2, \ p=0.18, \ I^2=42\% \)). Only one study (Dai et al \textit{42}) demonstrated a non-significant result, and it included the least proportion (4.5%) of IPAF subjects with a radiological and/or pathological UIP pattern in contrast to 54.6% and 36.7% in the other two studies (Oldham et al \textit{38} and Kim et al \textit{48}), respectively.

Figure 6  Forrest plot of the effect of percentage of predicted forced vital capacity (%FVC) on all-cause mortality of interstitial pneumonia with autoimmune features (IPAF). The effect of %FVC on all-cause mortality was reported in three studies, in which a total of 351 participants with 158 men (45.0%) were included. Meta-analysis of these studies demonstrated that %FVC was significantly associated with all-cause mortality of IPAF with an HR of 0.89 (95% CI 0.81 to 0.98, \( p=0.02 \)) with no heterogeneity.

Figure 7  Forrest plot of the effect of percentage of diffusing capacity of the lung for carbon monoxide (%DLCO) on all-cause mortality of interstitial pneumonia with autoimmune features (IPAF). The effect of %DLCO on all-cause mortality was reported in three studies, in which a total of 351 participants with 158 men (45.0%) were included. Meta-analysis of these studies demonstrated that %DLCO was significantly associated with all-cause mortality of IPAF with an HR of 0.77 (95% CI 0.69 to 0.85, \( p<0.00001 \)) with no heterogeneity.

a prodromal condition that could evolve into an overt CTD-IP or a new clinical entity that is different from CTD-IP or IIPs. If the former is the case, it is expected that IPAF-UIP would demonstrate a better prognosis than IPF because the prognosis of CTD-UIP was reported to be better than that of IPF.\textit{50} However, other reports also stated that rheumatoid arthritis-associated UIP (RA-UIP) would follow a similar prognosis as IPF.\textit{51} Therefore, the finding of this review that there was no significant difference of all-cause mortality between IPAF-UIP and IPF may be derived from the possibility that IPAF-UIP was a prodromal state for RA-UIP. In fact, IPAF was followed up to reveal the CTD incidence in two studies in this review, and the largest proportion of incident cases was demonstrated to be RA (58.3% and 30.8%, respectively; data not shown). However, this remains a speculation because no other studies have reported the development of defined CTDs. In addition, clinical information specific to RA such as anti-cyclic citrullinated peptide antibody\textit{52} was lacking. If the IPAF-UIP included in this review was positive for this autoantibody, it would likely have suggested underlying RA-UIP and reasonably explained the similar
Table 5
Assessment of quality of evidence of a prognostic factor for interstitial pneumonia with autoimmune features by the Grades of Recommendation, Assessment, Development and Evaluation (GRADE) system

| Outcomes | Prognostic factors | GRADE factors |
|----------|--------------------|--------------|
| All-cause mortality | Age | Phase |
| | Uni | 1 |
| | Multi | 1 |
| | Inconsistency | Study limitations |
| | Phase | 1 |
| | Study limitations | + |
| | Dose effect size | Imprecision |
| | Dose effect size | + |
| | Imprecision | + |
| | Publication bias | Indirectness |
| | + | + |
| | Indirectness | + |
| | Publication bias | + |
| | Overall quality | | |
| | | Very low |
| | | Very low |

This systematic review clarified a prognostic factor for IPAF. Age was significantly associated with worse all-cause mortality of the disease in both univariate and multivariate analyses. This finding is consistent with previous reports of both IPF and other types of IP and can be explained by the fact that the number of comorbidities increases and pulmonary function deteriorates with age. Although sex, smoking habit, radiological and/or pathological UIP pattern, %FVC and %DLCO were not confirmed as prognostic factors due to non-significant or inconsistent results or a significant result in a single study by multivariate analyses, it is not pathophysiologically unreasonable to believe that these factors might have prognostic significance. Smoking habit is common for men and increases the risk of other comorbidities such as hypertension and other cardiovascular diseases, which can affect the overall mortality of the disease. Smoking may also cause emphysematous changes in pulmonary parenchyma combined with pulmonary fibrosis, making prognosis worse. The presence of a radiological and/or pathological UIP pattern represents a progressive fibrotic process that is unfavourably responsive to traditional immunosuppressants although the prognostic difference between UIP and other IP patterns is not as obvious for CTD-IP as for IIPs. Furthermore, %FVC and %DLCO reflect the residual capacity of the lung, which may impact the prognosis of the disease.

The findings of prognostic factors in this review should be interpreted with caution. In particular, attention should be paid to the fact that significance was determined based on the results of multivariate analyses, which were conducted in only a few studies. As a result, some potential prognostic factors may possibly have been deemed as non-prognostic due to type 2 error in this review. Furthermore, any clinical information that was reported in fewer than three studies was excluded from further investigation for prognostic significance by assuming less clinical relevance. However, some clinical features of IPAF such as the positivity of clinical, serological and morphological domains still seem relevant to the prognosis of the disease.

mortality as IPF. Further research is imperative to clarify the prognostic significance of a UIP pattern for IPAF compared with IPF and what IPAF-UIP represents.

Another important factor that may have affected the prognosis of IPAF is that cases of IPAF in this review may not necessarily represent an inception cohort because all studies were retrospectively conducted at a referral centre. As a result, some cases of IPAF with gradual progression may have been referred and enrolled for further investigation at an advanced stage, and this would have led to a biased estimation of survival. In fact, only three studies described the follow-up period for IPAF, and it was reported as 1.5 years in one study. This period seems to be too short to estimate the prognosis of this disease group unless the rapidly progressive type of IP or acute exacerbation of the disease was the main subject of study. Therefore, it is important to enrol an inception cohort of IPAF to elucidate its prognosis in future research.
Further research needs to be conducted to address this issue. There were also other methodological limitations that affected the findings of this review. First, a subgroup analysis was not feasible due to a small number of studies, although there was clinical and methodological diversity between studies. The prognosis and prognostic factors of IPAF may probably be different depending on these various clinical features. However, at least the effect of age is likely to be true regardless of clinical diversity of IPAF because this finding was consistent between studies and no heterogeneity was identified in the meta-analysis. Second, all studies included in this review were affected by some risk of bias, and thus the quality of presented evidence was deemed as very low. Therefore, more primary studies of high quality need to be conducted to clarify the prognosis and prognostic factors of IPAF. Moreover, the current systematic review and meta-analysis should be updated with additional reports in sufficient numbers, and other outcomes of clinical relevance that were not addressed in this study, such as a progression of the disease and CTD incidence, should also be addressed in the future.

CONCLUSIONS
This systematic review and meta-analysis demonstrated that all-cause mortality of IPAF-UIP was not significantly different from that of IPF, although the prognosis of IPAF may be better if the proportion of a UIP pattern is small. Old age was significantly associated with worse all-cause mortality of IPAF.

Acknowledgements The authors would like to thank Cambridge English Correction Service for their professional proofreading.

Contributors HK planned the entire research project and analysed data. He also summarized the result and wrote the manuscript. He has full access to all of the data in the study and takes responsibility for the integrity of the data and the accuracy of the analysis. OM made contributions in conceiving the research project. He was also engaged in literature search, data extraction and risk-of-bias assessment. He provided thoughts and opinions to a draft paper and approved of the final version of the manuscript.

Funding The authors have not declared a specific grant for this research from any funding agency in the public, commercial or not-for-profit sectors.

Competing interests None declared.

Patient consent for publication Not required.

Ethics approval It was waived as this study was only based on the summary result of previously published articles and individual patient data were not obtained or accessed.

Provenance and peer review Not commissioned; externally peer reviewed.

Data availability statement Data are available upon reasonable request.

Open access This is an open access article distributed in accordance with the Creative Commons Attribution Non Commercial (CC BY-NC 4.0) license, which permits others to distribute, remix, adapt, build upon this work non-commercially, and license their derivative works on different terms, provided the original work is properly cited, appropriate credit is given, any changes made indicated, and the use is non-commercial. See: http://creativecommons.org/licenses/by-nc/4.0/.

ORCID iD
Hiroyuki Kamiya http://orcid.org/0000-0001-5623-1279

REFERENCES

1. Travis WD, Costabel U, Hansell DM, et al. An official American Thoracic Society/European Respiratory Society statement: update of the international multidisciplinary classification of the idiopathic interstitial pneumonias. Am J Respir Crit Care Med 2013;188:733–48.
2. Fischer A, du Bois R. Interstitial lung disease in connective tissue disorders. The Lancet 2012;380:688–98.
3. Gulati M, Redlich CA. Asbestosis and environmental causes of usual interstitial pneumonia. Curr Opin Pulm Med 2015;21:193–200.
4. Skorupski S, Weatherley N, Swift A, et al. Drug-induced interstitial lung disease: a systematic review. JCM 2018;7.
5. Park JH, Kim DS, Park IN, et al. Prognosis of fibrotic interstitial pneumonia: idiopathic versus collagen vascular disease-related subtypes. Am J Respir Crit Care Med 2007;175:705–11.
6. Tzelepis GE, Toya SP, Moutsopoulos HM. Occult connective tissue diseases mimicking idiopathic interstitial pneumonias. Eur Respir J 2008;31:11–20.
7. Homma Y, Ohtsuka Y, Tanimura K, et al. Can idiopathic pneumonia as the sole presentation of collagen vascular diseases be differentiated from idiopathic interstitial pneumonias? Chest 1995;108:246–51.
8. Fischer A, Pfazgraf FJ, Feghali-Bostwick CA, et al. Anti-th/τ- positivity in a cohort of patients with idiopathic pulmonary fibrosis. J Rheumatol 2006;33:1600–5.
9. Mosca M, Neri R, Bombardieri S. Undifferentiated connective tissue diseases (UCTD): a review of the literature and a proposal for preliminary classification criteria. Clin Exp Rheumatol 1999;17:615–20.
10. Suda T, Kono M, Nakamura Y, et al. Distinct prognosis of idiopathic nonspecific interstitial pneumonia (NSIP) fulfilling criteria for undifferentiated connective tissue disease (UCTD). Respir Med 2010;104:1527–34.
11. Kinder BW, Collard HR, Koth L, et al. Idiopathic nonspecific interstitial pneumonia: lung manifestation of undifferentiated connective tissue disease? Am J Respir Crit Care Med 2007;176:691–7.
12. Corte TJ, Copley SJ, Desai SR, et al. Significance of connective tissue disease features in idiopathic pulmonary fibrosis. Eur Respir J 2012;39:661–8.
13. Fischer A, West SG, Swigris JJ, et al. Connective tissue disease-associated interstitial lung disease: a call for clarification. Chest 2010;138:251–6.
14. Vij R, Noth I, Streek ME. Autoimmune-featured interstitial lung disease: a distinct entity. Chest 2011;140:1292–9.
15. Fischer A, Antoniou KM, Brown KK, et al. An official European Respiratory Society/American Thoracic Society research statement: interstitial pneumonia with autoimmune features. Eur Respir J 2015;46:976–87.
16. Chartrand S, Swigris JJ, Stanchev L, et al. Clinical features and natural history of interstitial pneumonia with autoimmune features: a single center experience. Respir Med 2016;119:150–4.
17. Sambataro G, Sambataro D, Torrisi SE, et al. State of the art in interstitial pneumonia with autoimmune features: a systematic review on retrospective studies and suggestions for further advances. Eur Respir Rev 2018;27.
18. Moher D, Liberati A, Tetzlaff J, et al. Preferred reporting items for systematic reviews and meta-analyses: the PRISMA statement. Ann Intern Med 2009;151:264–9.
19. Stroup DF, Berlin JA, Morton SC, et al. Meta-analysis of observational studies in epidemiology: a proposal for reporting. Meta-analysis of Observational Studies in Epidemiology (MOOSE) group. JAMA 2000;283:2008–12.
20. Aletaha D, Neogi T, Silman AJ, et al. Rheumatoid arthritis classification criteria: an American College of Rheumatology/European League Against Rheumatism collaborative initiative. Ann Rheum Dis 2010;2010:1580–8.
21. van den Hoogen F, Kallberg P, Fransen J, et al. Classification criteria for systemic sclerosis: an American College of Rheumatology/European League Against Rheumatism collaborative initiative. Arthritis Rheum 2011;64:1567–83.
22. Tan EM, Cohen AS, Fries JF, et al. The 1982 revised criteria for the classification of systemic lupus erythematosus. Arthritis Rheum 1982;25:1271–7.
23. Vitali C, Bombardieri S, Jonsson R, et al. Classification criteria for Sjögren's syndrome: a revised version of the European criteria proposed by the American-European consensus group. Ann Rheum Dis 2002;61:554–8.
24. Smolen JS, Steinert G. Mixed connective tissue disease: to be or not to be? Arthritis Rheum 1998;41:766–77.
25. Bohan A, Peter JB. Polymyositis and dermatomyositis (first of two parts). New Engl J Med 1975;292:344–7.
