The need for speed:
Maximizing the speed of random walk in fixed environments

Eviatar B. Procaccia*, Ron Rosenthal †
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Abstract
We study nearest neighbor random walks in fixed environments of \( \mathbb{Z} \) composed of two point types: \((\frac{1}{2}, \frac{1}{2})\) and \((p, 1-p)\) for \( p > \frac{1}{2} \). We show that for every environment with density of \( p \) drifts bounded by \( \lambda \) we have \( \lim \sup_{n \to \infty} \frac{X_n}{n} \leq (2p - 1)\lambda \), where \( X_n \) is a random walk in the environment. In addition up to some integer effect the environment which gives the greatest speed is given by equally spaced drifts.

1 Introduction
The subject of random walks in non-homogeneous environments received much interest in recent decades. There has been tremendous progress in the study of such random walks in a random environment, however not much is known about random walks in a given fixed environment. In this paper we study the maximal speed a nearest neighbor random walk can achieve while walking over \( \mathbb{Z} \), in a fixed environment composed of two types of drifts, \((p, 1-p)\) (i.e. probability \( p \) to go to the right, and probability \( 1-p \) to go to the left) and \((\frac{1}{2}, \frac{1}{2})\).

A similar question in the continuous setting was posed by Itai Benjamini and answered by Susan Lee. In [4] Lee proves that a diffusion process \( dX_t = b(X_t)dt + dB_t \) on the interval \([0, 1]\), with 0 as a reflecting boundary, \( b(x) \geq 0 \), and \( \int_0^1 b(x)dx = 1 \), has a unique \( b \) which minimize the expected time for \( X_t \) to hit 1, given by the step function \( 2 \cdot 1_{[1/4, 3/4]} \). This result is different in nature from the one we get for the discrete case as the optimal environment in our case is given by equally spaced drifts along \( \mathbb{Z} \). Notice that a major difference between Lee’s setup and the one in this paper is that in the later the diffusion coefficient and drift are coupled. Another problem similar in spirit is presented in [4], however the technical details are completely different. A related question for perturbation of simple random walk by a random environment of asymptotically small drifts, for which the recurrence/transience question becomes more involved is studied in [6].

The question of this paper arose while the first author and Noam Berger tried to give a speed bound for a non Markovian random walks over \( \mathbb{Z} \) and the application will be published in [2].

*Weizmann Institute of Science
†The Hebrew University of Jerusalem
In order to state the theorem we give a more precise definition of the environments we study:

**Definition 1.1.** Given $\frac{1}{2} < p \leq 1$ and $0 \leq \lambda \leq 1$ we call $\omega : \mathbb{Z} \to [0, 1]$ a $(p, \lambda)$ environment if the following holds:

1. For every $x \in \mathbb{Z}$ either $\omega(x) = \frac{1}{2}$ or $\omega(x) = p$.
2. 
   $$\limsup_{n \to \infty} \frac{1}{n+1} \sum_{x=0}^{n} 1_{\omega(x)=p} = \lambda.$$  \hspace{0.5cm} (1.1)

Throughout this paper we denote by $\{X_n\}_{n=0}^{\infty}$ a random walk on $\mathbb{Z}$ (or sub interval of it). In addition for a given environment $\omega : \mathbb{Z} \to [0, 1]$ and a point $x \in \mathbb{Z}$ we denote by $P^x_\omega$ the law of the random walk, which makes it into a stationary Markov chain with the following transition probabilities

$$P^x_\omega (X_{n+1} = y | X_n = z) = \begin{cases} \omega(z) & y = z + 1 \\ 1 - \omega(z) & y = z - 1 \\ 0 & \text{otherwise} \end{cases},$$

and initial distribution

$$P^x_\omega (X_0 = x) = 1.$$

The goal of this paper is to study the maximal speed a random walk in $(p, \lambda)$ environments can achieve, i.e. the behavior of the random variable $\limsup_{n \to \infty} \frac{X_n}{n}$.

We start with a simple observation regarding the random variable $\limsup_{n \to \infty} \frac{X_n}{n}$:

**Lemma 1.2.** For every $(p, \lambda)$ environment $\omega$ and every $x \in \mathbb{Z}$ the random variable $\limsup_{n \to \infty} \frac{X_n}{n}$ is a $P^x_\omega$ almost sure constant.

The main theorem we prove is an upper bound on the speed of random walks in $(p, \lambda)$ environments:

**Theorem 1.3.** For every $(p, \lambda)$ environment $\omega$ and every $x \in \mathbb{Z}$

$$\limsup_{n \to \infty} \frac{X_n}{n} \leq (2p - 1)\lambda, \quad P^x_\omega \text{ a.s.}$$

As a result from the theorem we have the following corollary for random walks in random environments (RWRE):

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Corollary 1.4. Let $P$ be a stationary and ergodic probability measure on environments $\omega$ of $\mathbb{Z}$ such that $P(\omega(0) = p) = \lambda$ and $P(\omega(0) = 1/2) = 1 - \lambda$ for some $0 \leq \lambda \leq 1$ and $1/2 < p \leq 1$. Let $\{X_n\}$ be a RWRE with environment $\omega$ distributed according to $P$ (for a more precise definition of RWRE see [8]), then

$$\lim_{n \to \infty} \frac{X_n}{n} \leq (2p - 1)\lambda,$$

for $P$ almost every environment $\omega$ and $P^\omega$ almost every random walk in it.

The main idea beyond the proof of Theorem 1.3 is an exact calculation of some expected hitting times in a finite segment with a particular environment. We show that the expected hitting time of a random walk starting at the origin and reflected there, to the point $N$, where there are $k$ drift points between the origin and $N$, can be described by

$$E^0_\omega[T_N] = \frac{N^2}{(2p - 1) \cdot k + 1} + \langle H_k(l - b), (l - b) \rangle,$$

where $l$ is the vector of drift positions, $b$ is a fixed vector and $H_k$ is a $k \times k$ symmetric positive definite matrix depending only on $p$. For the full proposition and definitions see Proposition 2.2.

The last equation implies a lower bound on $T_N$ and hence, eventually, an upper bound on the speed.

A natural question that arises is whether the inequality of Theorem 1.3 can be improved. In section 5 we prove the following results:

Proposition 1.5. Let $m \in \mathbb{N}$ and let $\lambda = \frac{1}{m}$. Let $\omega$ be an environment s.t $\omega(i \cdot m) = p$, $\forall i \in \mathbb{Z}$, and $\omega(x) = \frac{1}{2}$, $\forall x \notin \{i \cdot m : i \in \mathbb{Z}\}$, then a random walk $\{X_n\}$ in $\omega$ has the property

$$\limsup_{n \to \infty} \frac{X_n}{n} = \lim_{n \to \infty} \frac{X_n}{n} = (2p - 1)\lambda.$$

Proposition 1.6. For every $p$ and $\lambda > 0$, there exists a $(p, \lambda)$ environment $\omega$, and a constant $D(p)$ such that

$$\limsup_{n \to \infty} \frac{X_n}{n} \geq (2p - 1)\lambda - D(p)\lambda^3.$$

We also show the lower bound in proposition 1.6 can not be improved for all values of $p$.

Proposition 1.7. Let $\lambda > 0$ be of the form $\lambda = \frac{n}{mn+l}$, such that $\lambda \neq \frac{1}{k}$, for all $k \in \mathbb{N}$. There exists a constant $D = D(n) > 0$ such that and every $(1, \lambda)$ environment $\omega$ we have $\limsup_{N \to \infty} \frac{X_N}{N} \leq \lambda - D\lambda^3$.

The structure of this paper is as follows: Section 2 deals with a particular finite case of the problem which stands in the heart of the proof of the infinite case. Section 3 contains the proof of theorem 1.3. Section 4 deals with RWRE. In section 5 we discuss tightness of the result. In section 6 we prove Lemma 1.2. Finally, in section 7 we give some conjectures and open questions regarding the model.
2 Finite environment with reflection at the origin

We start by analyzing a finite variant of the problem. Consider nearest neighbor random walks on subsets of \( \mathbb{Z} \) of the form \( \{0, 1, \ldots, N\} \), with reflection at the origin, an absorbing state at \( N \), and the rest of the points are either \( (\frac{1}{2}, \frac{1}{2}) \) or \( (p, 1 - p) \). More precisely we study the following environments:

**Definition 2.1.** Given \( N \in \mathbb{N}, \frac{1}{2} < p \leq 1 \) and \( k \in \mathbb{N} \) such that \( 1 \leq k \leq N - 1 \), we call \( \omega : \{0, 1, \ldots, N - 1\} \rightarrow [0, 1] \) a \((N, p, k)\) environment on \( \{0, 1, \ldots, N\} \) if there exists \( L = \{l_i\}_{i=1}^k \subset \mathbb{N} \) such that

\[
0 < l_1 < l_2 < \ldots < l_k < N
\]

and

\[
\omega(x) = \begin{cases} 
1 & x = 0 \\
p & x \in L \\
\frac{1}{2} & x \in \{1, \ldots, N - 1\} \setminus L
\end{cases}
\]

Throughout this section \( T_N \) will denote the first time a random walk in a \((N, p, k)\) environment \( \omega \) hits \( N \), i.e, \( T_N = \min\{n \geq 0 : X_n = N\} \). In addition we use the following notations:

1. \( [N] = \{0, 1, \ldots, N\} \)
2. \( l = (l_1, \ldots, l_k) \)

and

3. \( l_0 = 0, \ l_{k+1} = N. \)

The following is the main proposition of this section:

**Proposition 2.2.** For every \((N, p, k)\) environment \( \omega \) we have

\[
E^0_\omega[T_N] \geq \frac{N^2}{(2p - 1)k + 1}. \tag{2.1}
\]

In addition there exists a \((N, p, k)\) environment which satisfies equality if and only if both \( \frac{(2p-1)N}{(2p-1)k+1} \) and \( \frac{pN}{(2p-1)k+1} \) are integers. Furthermore there exists a \( k \times k \) positive definite symmetric matrix \( H_k \), with entries depending only on \( p \), such that

\[
E^0_\omega[T_N] = \frac{N^2}{(2p - 1)k + 1} + \langle H_k(l - b), (l - b) \rangle, \tag{2.2}
\]

where \( \langle ., . \rangle \) denotes the standard inner product, and \( b = (b_1, \ldots, b_k) \) is the vector given by

\[
b_i = \frac{(2p - 1)i + (1 - p)}{(2p - 1)k + 1} N. \tag{2.3}
\]
Proof. Define \( v : [N] \to \mathbb{R} \) by \( v(x) = E^x_\omega[T_N] \). By conditioning on the first step and using linearity of the expectation one observes that \( v \) satisfies the following equations:

\[
v(0) = v(1) + 1 \\
v(1) = 0 \\
v(x) = \frac{1}{2}v(x + 1) + \frac{1}{2}v(x - 1) + 1, \quad \forall x \in \{1, 2, \ldots, N - 1\} \setminus L \\
v(x) = p \cdot v(x + 1) + (1 - p) \cdot v(x - 1) + 1, \quad \forall x \in L.
\] (2.4)

Restricting ourselves to an interval of the form \([l_{j-1}, l_j]\), for some \( 1 \leq j \leq N \), we see that the solution to the equations

\[
v(x) = \frac{1}{2}v(x + 1) + \frac{1}{2}v(x - 1) + 1, \quad \forall l_{j-1} < x < l_j,
\]
is given by \( v(x) = -x^2 + C_j \cdot x + D_j \) with \( C_j \) and \( D_j \) two constants determined by the value of \( v \) at \( x = l_{j-1} \) and \( x = l_j \). Thus one can replace the equations in (2.4) with the following ones:

\[
v(0) = v(1) + 1 \\
v(N) = 0 \\
v(x) = -x^2 + C_j x + D_j, \quad \forall x \in [l_{j-1}, l_j] \quad \forall 1 \leq j \leq k + 1 \\
v(l_j) = p \cdot v(l_j + 1) + (1 - p) \cdot v(l_j - 1) + 1, \quad \forall 1 \leq j \leq k.
\] (2.5)

Solving those equations one finds that

\[
C_1 = 0 \\
C_j = \frac{2(2p - 1)}{1 - p} \sum_{i=1}^{j-1} \left( \frac{1 - p}{p} \right)^{j-i} l_i, \quad \text{for } 2 \leq j \leq k + 1
\]

\[
D_{k+1} = N^2 - N \cdot \frac{2(2p - 1)}{1 - p} \sum_{i=1}^{k} \left( \frac{1 - p}{p} \right)^{k+1-i} l_i \\
D_j = N^2 - \frac{2(2p - 1)}{p} N \sum_{i=1}^{k} \left( \frac{1 - p}{p} \right)^{k-i} l_i \\
+ \frac{2(2p - 1)}{p} \sum_{i=j}^{k} l_i^2 - \frac{2(2p - 1)^2}{p^2} \sum_{i=j}^{k} \sum_{m=1}^{i-1} \left( \frac{1 - p}{p} \right)^{i-m-1} l_i l_m, \quad \text{for } 1 \leq j \leq k.
\] (2.6)

In particular we get that

\[
f(l_1, \ldots, l_k) := E^0_\omega[T_N] = D_1 = N^2 - \frac{2(2p - 1)}{p} N \sum_{i=1}^{k} \left( \frac{1 - p}{p} \right)^{k-i} l_i \\
+ \frac{2(2p - 1)}{p} \sum_{i=1}^{k} l_i^2 - \frac{2(2p - 1)^2}{p^2} \sum_{i=1}^{k} \sum_{m=1}^{i-1} \left( \frac{1 - p}{p} \right)^{i-m-1} l_i l_m.
\] (2.7)

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Notice that the last function is a polynomial of degree two in \( l_1, \ldots, l_k \).

One can check by substitution that the vector \( b = (b_1, \ldots, b_k) \), defined in \( (2.3) \), is a solution to the equation \( \text{grad} f(l) = 0 \), which makes \( b \) into an extremum point of \( f \). In addition the Hessian of \( f \) is constant (not depending on \( N \) or \( l_1, \ldots, l_k \)) and is given by the matrix

\[
H_k = -\frac{2(2p - 1)^2}{p^2} \cdot \begin{pmatrix}
-\frac{2p}{2p-1} & 1 & \left(\frac{1-p}{p}\right) & \left(\frac{1-p}{p}\right)^2 & \cdots & \cdots & \left(\frac{1-p}{p}\right)^{k-2} \\
1 & -\frac{2p}{2p-1} & 1 & \left(\frac{1-p}{p}\right) & \left(\frac{1-p}{p}\right)^2 & \cdots & \left(\frac{1-p}{p}\right)^{k-3} \\
\left(\frac{1-p}{p}\right) & 1 & -\frac{2p}{2p-1} & 1 & \cdots & \cdots & \left(\frac{1-p}{p}\right)^{k-4} \\
\left(\frac{1-p}{p}\right)^2 & \left(\frac{1-p}{p}\right) & 1 & \ddots & \ddots & \ddots & \\
\vdots & \vdots & \ddots & \ddots & \ddots & \ddots & \ddots \\
\left(\frac{1-p}{p}\right)^{k-2} & \left(\frac{1-p}{p}\right)^{k-1} & \cdots & 1 & -\frac{2p}{2p-1}
\end{pmatrix}
\]

We also define the matrix \( M_k \) by

\[
H_k \equiv -\frac{2(2p - 1)^2}{p^2} \cdot M_k.
\]

We notice that for \( 1 \leq j \leq k \), the \( j \)th principal minors of \( H_k \) and \( M_k \) are exactly \( H_j \) and \( M_j \) respectively.

By subtracting the \((k-1)\)th column and row multiplied by \( \frac{1-p}{p} \) of \( M_k \) from the \( k \)th column and row respectively, one gets the following recursion formula for the determinant of \( M_k \):

\[
det(M_k) = \left( -\frac{2p}{2p-1} - \frac{2(1-p)}{p} - \frac{2(1-p)^2}{p(2p-1)} \right) \det(M_{k-1}) - \left( 1 + \frac{2(1-p)}{2p-1} \right)^2 \det(M_{k-2}).
\]

Therefore, using induction one gets

\[
det(M_k) = (-1)^k \cdot \frac{k(2p - 1) + 1}{(2p - 1)^k} \cdot \frac{(2p - 1)^2}{p^{2k}}.
\]

Since \( \det(H_k) \) is positive for every \( \frac{1}{2} < p \leq 1 \) and \( k \in \mathbb{N} \), it follows by Silvester’s criterion (see [3]) that \( H_k \) is a positive definite matrix, and therefore \( b \) is the unique absolute minimum of \( f = E^0_\omega[T_N] \). Finally, by rearranging \( f \) one can show that

\[
f(l) = E^0_\omega[T_N] = \frac{N^2}{(2p-1)k + 1} + \langle H_k(l-b), (l-b) \rangle.
\]

From the last formula we get \( E^0_\omega[T_N] \geq \frac{N^2}{(2p-1)k+1} \) and equality holds if and only if \( l = b \). One can see from the definition of \( b \) that such \( l \) defines a \((N, p, k)\) environment if and only if both \( \frac{(2p-1)N}{(2p-1)k+1} \) and \( \frac{pN}{(2p-1)k+1} \) are integers. \( \square \)
Before turning to the infinite case we give a uniform bound on the norm of the matrices $H_k$, which will be used in Section 5.

**Lemma 2.3.** There exists some finite positive constant $C = C(p)$ such that

$$\sup_{k \in \mathbb{N}} \|H_k\|_2 \leq C.$$

**Proof.** Fix $k \in \mathbb{N}$ and for $1 \leq i \leq k$ denote by $r_k(i), c_k(i)$ the $i^{th}$ row and column of the matrix $H_k$ respectively. We notice that

$$\|r_k(i)\|_1 = \frac{2(2p-1)^2}{p^2} \cdot \left( \frac{2p}{2p-1} + \sum_{j=0}^{k-1-i} \left( \frac{1-p}{p} \right)^j + \sum_{j=0}^{i-2} \left( \frac{1-p}{p} \right)^j \right),$$

and

$$\|H_k\|_2 \leq \sqrt{\|H_k\|_1 \cdot \|H_k\|_\infty},$$

we get that for every $k \in \mathbb{N}$

$$\|H_k\|_2 \leq C(p).$$

**3 Proof of the main theorem**

Fix $\frac{1}{2} < p \leq 1$ and $0 \leq \lambda \leq 1$. We start with the following estimation of $\mathbf{E}_\omega[T_N]$:
Lemma 3.1. Given two $\mathbb{Z}$ environments $\omega, \bar{\omega}$ such that for every $x \in \mathbb{Z}$, $\omega(x) \leq \bar{\omega}(x)$. Denote by $T_n, \bar{T}_n$ the hitting times in the environments $\omega, \bar{\omega}$ respectively, then for every $n > 0$, $T_n$ stochastically dominates $\bar{T}_n$, i.e $P^0_\omega(T_n > t) \geq P^0_{\bar{\omega}}(\bar{T}_n > t)$.

Proof. This lemma follows from a standard coupling argument. Let $U_n \sim U[0, 1]$ be a sequence of i.i.d random variables. Let $P_{\omega, \bar{\omega}}$ be the joint measure of two processes $X_n$ and $\bar{X}_n$ such that both the processes at time $n$ move according to $U_n$ and the environments $\omega$ and $\bar{\omega}$, i.e.

$$P_{\omega, \bar{\omega}}(X_{n+1} = x \pm 1, \bar{X}_{n+1} = \bar{x} \pm 1 | X_n = x, \bar{X}_n = \bar{x}) = P_{\omega, \bar{\omega}}(U_n \leq \omega(x), U_n \leq \bar{\omega}(\bar{x})),$$  \hspace{1cm} (3.1)

and

$$P_{\omega, \bar{\omega}}(X_0 = 0, \bar{X}_0 = 0) = 1.$$  \hspace{1cm} (3.2)

By this coupling whenever the processes meet at some point, the random walk $\bar{X}_n$ has a higher probability to turn right. We therefore obtain that $P_{\omega, \bar{\omega}}$ a.s for every $n \in \mathbb{N}$, $\bar{X}_n \geq X_n$, thus $P_{\omega, \bar{\omega}}$ a.s $\bar{T}_n \leq T_n$. \hfill \Box

We turn now to prove the main theorem.

Proof of Theorem 1.3. Let $\epsilon > 0$, and let $\omega$ be a $(p, \lambda)$ environment. Since $\omega$ is a $(p, \lambda)$ environment there exists $M \in \mathbb{N}$ such that for every $N \geq M$ we have

$$\frac{\# \{x \in [N] : \omega(x) = p \}}{N} \leq \lambda + \epsilon.$$  \hspace{1cm} (3.3)

For $N \geq M$ we define a new environment $\bar{\omega}$ as follows :

$$\bar{\omega}(x) = \begin{cases} \omega(x) & N \nmid x \\ 1 & N \mid x, \end{cases}$$

where $N \mid x$ is a shorthand for $N$ divides $x$.

Let $\bar{T}_n$ be the same hitting time distributed according to the environment $\bar{\omega}$. Since for every $x \in \mathbb{Z}$ we have $\omega(x) \leq \bar{\omega}(x)$ it follows, using Lemma 3.1, that

$$\frac{T_{nN}}{nN} = \frac{1}{n} \sum_{k=1}^{n} \frac{T_{kN} - T_{(k-1)N}}{N} \geq \frac{1}{n} \sum_{k=1}^{n} \frac{T_{kN} - T_{(k-1)N}}{N}.$$  \hspace{1cm} (3.4)

By the strong Markov property the random variables $\{\bar{T}_{kN} - \bar{T}_{(k-1)N}\}_{k=1}^\infty$ are independent (but for general environment not identically distributed) and we wish to apply Kolmogorov’s strong law of large numbers.

For $n \in \mathbb{N}$ denote by $S_n$ the first hitting time of $n$ by a symmetric simple random walk with reflection at the origin and starting at 0. By Lemma 3.1 for every $k \in \mathbb{N}$ we have that $S_N$ stochastically dominates $T_{kN} - \bar{T}_{(k-1)N}$, and therefore

$$E^{\omega}_n \left[ \frac{T_{kN} - T_{(k-1)N}}{N} \right] \leq E^{0} \left[ \frac{S_N}{N} \right] = N < \infty,$$  \hspace{1cm} (3.4)
and
\[ E^0 \left[ \left( \frac{T_{kN} - T_{(k-1)N}}{N} \right)^2 \right] \leq E^0 \left[ \left( \frac{S_N}{N} \right)^2 \right] \leq \frac{5}{3} N^2 < \infty. \]

The last relations are derived from the optional stopping theorem (see [7] Theorem 12.20) and the fact that for a symmetric simple random walk \( Y_n \),
\[ Y_n^2 - n, \]
\[ Y_n^4 - 6nY_n^2 + 3n^2 + 2n \]
are martingales. It therefore follows by Kolmogorov’s strong law of large numbers that
\[ \lim_{n \to \infty} \frac{1}{n} \sum_{k=1}^{n} \frac{T_{kN} - T_{(k-1)N}}{N} - E \left[ \frac{1}{n} \sum_{k=1}^{n} \frac{T_{kN} - T_{(k-1)N}}{N} \right] = 0, \ P_0^\omega \ a.s. \quad (3.5) \]

For \( 1 \leq k \leq n \) we define \( \lambda_k \) to be the \( p \)'s density in the interval \( [(k-1)N, kN) \), i.e.
\[ \lambda_k = \frac{1}{N} \sum_{x=(k-1)N}^{kN-1} \mathbb{1}_{\omega(x)=p}. \]

By (3.6) we have
\[ \frac{1}{n} \sum_{k=1}^{n} \lambda_k = \frac{1}{n} \sum_{k=1}^{n} \frac{1}{N} \sum_{x=(k-1)N}^{kN-1} \mathbb{1}_{\omega(x)=p} = \frac{\#\{x \in [nN - 1] : \omega(x) = p\}}{nN} < \lambda + \epsilon. \quad (3.6) \]

Notice that each of the segments \( [(k-1)N, kN - 1] \) of \( \bar{\omega} \) is a \( (N, p, \lambda_k N) \) environment. It therefore follows by Proposition 2.2 that
\[ \frac{1}{n} \sum_{k=1}^{n} E^\omega \left[ \frac{T_{kN} - T_{(k-1)N}}{N} \right] \geq \frac{1}{n} \sum_{k=1}^{n} \frac{1}{(2p-1)} \cdot \lambda_k + \frac{1}{N} \geq \frac{\sum_{k=1}^{n} (2p-1) \cdot \lambda_k + \frac{1}{N}}{nN} \]
\[ \geq \frac{1}{(2p-1)(\lambda + \epsilon) + \frac{1}{nN}}, \quad (3.7) \]

where the second inequality follows from the inequality of arithmetic and harmonic means and the third is by (3.6). Thus,
\[ \liminf_{n \to \infty} \frac{T_n}{n} \overset{P_0^\omega, \ a.s.}{\geq} \liminf_{n \to \infty} \frac{1}{n} \sum_{k=1}^{n} \frac{T_{kN} - T_{(k-1)N}}{N} = \liminf_{n \to \infty} \frac{1}{n} \sum_{k=1}^{n} E \left[ \frac{T_{kN} - T_{(k-1)N}}{N} \right] \]
\[ \geq \lim_{n \to \infty} \frac{1}{(2p-1)(\lambda + \epsilon) + \frac{1}{nN}} = \frac{1}{(2p-1)(\lambda + \epsilon)}. \quad (3.8) \]

Since \( \epsilon > 0 \) was arbitrary we obtain
\[ \liminf_{n \to \infty} \frac{T_n}{n} \geq \frac{1}{(2p-1)\lambda}. \]

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with the notation \( \frac{1}{0} = \infty \). Now for \( n \in \mathbb{N} \) let \( k_n \) be the unique random integers such that \( T_{k_n} \leq n < T_{k_n+1} \).

Since \( X_n < k_n + 1 \) we get that

\[
\frac{X_n}{n} - \frac{1}{n} \leq \frac{k_n}{n}.
\]

Thus

\[
\limsup_{n \to \infty} \frac{X_n}{n} \leq \limsup_{n \to \infty} \frac{k_n}{n} \leq \limsup_{n \to \infty} \frac{n}{T_n} = \frac{1}{\liminf_{n \to \infty} \frac{T_n}{n}} \leq (2p - 1)\lambda. \tag{3.10}
\]

4 Application of the result to RWRE

We turn now to prove corollary 1.4.

**Proof of Corollary 1.4.** By ergodicity we obtain

\[
\lim_{n \to \infty} \frac{1}{n} \sum_{x=0}^{n-1} \mathbb{1}_{\omega(x) = p} = \lambda, \quad P \text{ a.s.} \tag{4.1}
\]

Define two random variables:

\[
\tilde{S} = \sum_{i=1}^{\infty} \frac{1}{\omega(-i)} \prod_{j=0}^{i-1} \rho(-j) + \frac{1}{\omega(0)},
\]

\[
\tilde{F} = \sum_{i=1}^{\infty} \frac{1}{(1 - \omega(i))} \prod_{j=0}^{i-1} \rho(j)^{-1} + \frac{1}{(1 - \omega(0))}, \tag{4.2}
\]

where \( \rho(j) = \frac{1 - \omega(j)}{\omega(j)} \). Since \( \forall i \in \mathbb{N}, \omega(i) \geq 1/2 \), it follows that \( \rho(i)^{-1} \geq 1 \) and therefore \( \tilde{F} = \infty, \) \( P \) a.s. By Lemmas 2.1.9 and 2.1.12 of [8], if \( E[\tilde{S}] = \infty \) and \( E[\tilde{F}] = \infty \) then \( \lim_{n \to \infty} \frac{X_n}{n} = 0, \) \( P^{0}_\omega \) a.s for \( P \) almost every \( \omega \), and if \( E[\tilde{S}] < \infty \) then \( \lim_{n \to \infty} \frac{X_n}{n} = E[T_1] \) and \( E[T_1] < \infty \), where \( E \) is the annealed expectation. By [8] Lemmas 2.1.10 and 2.1.12 we have, \( \lim_{n \to \infty} \frac{T_n}{n} \overset{a.s.}{=} E[T_1], \{T_{i+1} - T_i\}_{i=0}^{\infty} \) is a stationary and ergodic sequence and \( \lim_{n \to \infty} \frac{T_n}{n} = \lim_{n \to \infty} \frac{1}{n} \sum_{i=0}^{n-1} E[T_{i+1} - T_i] = E[T_1]. \) Let \( \lambda_n \) be the density of drifts in the interval \([0, n-1]\). By Proposition 2.2 and (4.1) we have

\[
E[T_1] = \lim_{n \to \infty} E \left[ \frac{T_n}{n} \right] \geq \lim_{n \to \infty} E \left[ \frac{n}{(2p-1)\lambda n - 1} \right] = \frac{1}{(2p-1)\lambda},
\]

and therefore

\[
\lim_{n \to \infty} \frac{X_n}{n} \leq (2p - 1)\lambda,
\]

for \( P \) almost every environment \( \omega \) and \( P^\omega \) almost every random walk in it. \( \square \)
It was pointed to the authors by Ofer Zeitouni that a trivial bound to the speed exists to RWRE and one needs to check that this bound is not better than the bound we get in Corollary 1.4. Note that the trivial bound is by no means tight. By $\lim_{n \to \infty} \frac{X_n}{n} = \frac{1}{E[S]}$. Thus in order to get an upper bound on the speed a lower bound on $E[S]$ is needed.

\[
E[S] = \sum_{i=1}^{\infty} E \left[ \frac{1}{\omega_i} \prod_{j=0}^{i-1} \rho_{-j} \right] + E \left[ \frac{1}{\omega_0} \right]
\]

\[
= \sum_{i=1}^{\infty} E \left[ e^{\sum_{j=0}^{i-1} \log \rho_{-j} - \log \omega_i} \right] + E \left[ \frac{1}{\omega_0} \right]
\]

\[
\geq \sum_{i=1}^{\infty} e^{\sum_{j=0}^{i-1} E[\log \rho_{-j}] - E[\log \omega_i]} + E \left[ \frac{1}{\omega_0} \right]
\]

\[
= p^{-\lambda} 2^{1-\lambda} \sum_{i=1}^{\infty} e^{\lambda \log \left( \frac{1-p}{p} \right)} + \frac{\lambda}{p} + (1 - \lambda) 2
\]

\[
= p^{-\lambda} 2^{1-\lambda} \frac{(1-p)^{\lambda}}{1 - \left( \frac{1-p}{p} \right)^{\lambda}} + \frac{\lambda}{p} + (1 - \lambda) 2, \quad (4.3)
\]

where the inequality is Jensen’s. Denote by $S(p, \lambda) = p^{-\lambda} 2^{1-\lambda} \frac{(1-p)^{\lambda}}{1 - \left( \frac{1-p}{p} \right)^{\lambda}} + \frac{\lambda}{p} + (1 - \lambda) 2$. In figure 4.1 we drew the difference between the bound from Corollary 1.4 and $S(p, \lambda)$. One can see that for a large region of $p$ and $\lambda$ the bound archived in Corollary 1.4 is tighter.

Figure 4.1: Difference of bounds
Remark 4.1. Note that in the case of i.i.d RWRE an explicit value of the speed can be calculated and is equal to \( \frac{(2p-1)\lambda}{\lambda+2p(1-\lambda)} \) which is always smaller than \((2p-1)\lambda\).

5 Tightness of the result

In this section we discuss tightness of the result in the sense: Is there a \((p, \lambda)\) environment \(\omega\), such that a random walk \(\{X_n\}\) in \(\omega\) has the property

\[
\limsup_{n \to \infty} \frac{X_n}{n} = \lim_{n \to \infty} \frac{X_n}{n} = (2p-1)\lambda, \quad P_\omega \text{ a.s.}
\]

5.1 Positive tightness

Proposition. Let \(m \in \mathbb{N}\) and assume \(\lambda = \frac{1}{m}\). Let \(\omega\) be an environment defined by

\[
\omega(x) = \begin{cases} 
    p & \text{if } x \in m\mathbb{Z} \\
    \frac{1}{2} & \text{otherwise} 
\end{cases}, 
\]

then a random walk \(\{X_n\}\) in \(\omega\) has the property

\[
\limsup_{n \to \infty} \frac{X_n}{n} = \lim_{n \to \infty} \frac{X_n}{n} = (2p-1)\lambda, \quad P_\omega \text{ a.s.}
\]

Proof. We prove this proposition by a direct calculation of the speed.

\[
P_\omega^0(T_m < T_{-m}) = p \left[ P_\omega^1(T_m < T_0) + P_\omega^1(T_m > T_0) P_\omega^0(T_m < T_{-m}) \right] \\
+ (1-p) P_\omega^{-1}(T_0 < T_{-m}) P_\omega^0(T_m < T_{-m}),
\]

but \(P_\omega^1(T_m < T_0) = P_\omega^{-1}(T_0 > T_{-m}) = \frac{1}{m}\), thus

\[
P_\omega^0(T_m < T_{-m}) = p.
\]

Now \(E_\omega[T_m \wedge T_{-m}] = m^2\). Consider \(\omega\) as an environment with a constant drift \(p\), such that every jump takes on average \(m^2\) steps. The speed of a random walk in an environment with constant drift \(p\) at any point is \((2p-1)\). Thus the speed for \(\omega\) is \((2p-1)\frac{m^2}{m^2} = (2p-1)\lambda\). \(\square\)

We turn to prove a general tightness result.

Proposition. For every \(\frac{1}{2} < p \leq 1\) and \(0 < \lambda \leq 1\), there exists a \((p, \lambda)\) environment \(\omega\), and a constant \(D(p) > 0\) such that

\[
\lim_{n \to \infty} \frac{X_n}{n} \geq (2p-1)\lambda - D(p)\lambda^3, \quad P_\omega \text{ a.s.}
\]
Proof. First assume that \( \lambda \in \mathbb{Q} \). We define the environment \( \omega \) by the positions \( \{l_i\} \) of non-zero drifts on \( \mathbb{Z} \). For every \( i \in \mathbb{Z} \) let \( l_i = \left\lfloor \frac{1}{\lambda} \left( i + \frac{1-p}{2p-1} \right) \right\rfloor \). Note that since \( \lambda \leq 1 \) all the drift positions are distinct, and \( \omega \) is indeed a \((p, \lambda)\) environment. For every \( N \in \mathbb{N} \) we denote by \( k = k(N) \), the number of drifts in the interval \([0, N]\). Note that \( \lim_{N \to \infty} \frac{k(N)}{N} = \lambda \). For a given \( k \in \mathbb{N} \) we denote \( b_{[k]} = (b_1, \ldots, b_k) \), where

\[
    b_i = \frac{(2p-1)i + (1-p)}{(2p-1)k + 1} N. \tag{5.5}
\]

By the Cauchy-Schwarz inequality

\[
    \langle H_k(l_{[k]} - b_{[k]}), (l_{[k]} - b_{[k]}) \rangle \leq \|H_k(l_{[k]} - b_{[k]})\|_2 \| (l_{[k]} - b_{[k]}) \|_2 
\leq \|H_k\|_2 \| (l_{[k]} - b_{[k]) \|_2^2. \tag{5.3}
\]

By Proposition 2.3 there exists a \( C(p) \) such that \( \|H_k\|_2 \leq C(p) \) for every \( k \in \mathbb{N} \). Thus

\[
    \lim_{N \to \infty} \frac{1}{N} \langle H_k(l_{[k]} - b_{[k]}), (l_{[k]} - b_{[k]}) \rangle \leq C(p) \lambda \lim_{k \to \infty} \frac{1}{k} \| (l_{[k]} - b_{[k]) \|_2^2.
\]

Notice that there exists a constant \( C' \) (does not depend on \( \lambda \) or any other parameter) such that for \( k \) large enough \( \| (l_{[k]} - b_{[k]}) \|_\infty \leq C' \), thus \( \| (l_{[k]} - b_{[k]}) \|_2 \leq C' \sqrt{k} \), therefore

\[
    \lim_{N \to \infty} \frac{1}{N} \langle H_k(l_{[k]} - b_{[k]}), (l_{[k]} - b_{[k]}) \rangle \leq C(p) C' \lambda. \tag{5.4}
\]

Next we prove that for the environment \( \omega \), the limit \( \lim_{n \to \infty} \frac{X_n}{n} \) exists. From Lemma 2.1.17 of [8] it is enough to show the limit \( \lim_{n \to \infty} \frac{T_n}{n} \) exists. Since \( \lambda \) is rational there exists some \( n_0 \in \mathbb{N} \) such that \( \frac{n_0}{\lambda} \) is an integer and therefore for every \( i \in \mathbb{N} \), \( l_{n_0 i} = \left\lfloor \frac{i n_0}{\lambda} + \frac{1-p}{\lambda(2p-1)} \right\rfloor = \frac{2n_0}{\lambda} + \left\lfloor \frac{1-p}{\lambda(2p-1)} \right\rfloor \).

It follows that \( \omega \) is \( n_0 \) periodic and therefore \( \{T_{kn} - T_{(k-1)n}\}_{k=2}^\infty \) are i.i.d. From the law of large numbers (Note that the first random variable in the sum is bounded and therefore negligible)

\[
    \lim_{k \to \infty} \frac{1}{k} \sum_{j=1}^{k} [T_{jn_0} - T_{(j-1)n_0}] = E[T_{n_0}], \quad P_\omega \text{ a.s.}
\]

Now define \( m_N \) to be the minimal integer such that \( m_N \cdot n_0 \leq N < (m_N + 1) n_0 \), then \( \frac{T_{m_N n_0}}{(m_N + 1)n_0} \leq \frac{T_N}{N} \leq \frac{T_{(m_N + 1)n_0}}{m_N n_0} \), thus

\[
    \lim_{N \to \infty} \frac{T_N}{N} = \lim_{k \to \infty} \frac{T_{kn_0}}{kn_0} = \frac{1}{n_0} E[T_{n_0}], \quad P_\omega \text{ a.s.} \tag{5.5}
\]

Note that if \( \lim_{n \to \infty} \frac{X_n}{n} \) exists, it is the same for the environment \( \omega \) and \( \bar{\omega} \) (\( \omega \) with reflection at the origin), since the random walk almost surely spends only a finite time left of the origin. By (5.5) the limit, \( \lim_{N \to \infty} \frac{X_N}{N} \) exists and by Lemma 2.1.1 of [8], \( \lim_{N \to \infty} \frac{X_N}{N} = \lim_{N \to \infty} \frac{N}{E[T_N]} \). We obtain from (5.4) and Proposition 2.2

\[
    \lim_{N \to \infty} \frac{X_N}{N} = \lim_{N \to \infty} \frac{N}{E[T_N]} \geq \frac{1}{(2p-1)\lambda + C' C(p) \lambda} \geq (2p-1)\lambda - C' C(p)(2p-1)^2 \lambda^3. \tag{5.6}
\]
Now For $\lambda \notin \mathbb{Q}$, let $\epsilon > 0$ and let $0 < \lambda' < 1$ be a rational number such that $\lambda - \epsilon < \lambda' < \lambda$. Define $\omega$ to be the environment defined above for the rational density $\lambda'$. Notice that $\omega$ is a $(p, \lambda')$ environment but also a $(p, \lambda)$ environment since $\lambda' < \lambda$. It follows from (5.6) that

$$
\lim_{N \to \infty} \frac{X_N}{N} \geq (2p - 1)\lambda' - C'C(p)(2p - 1)^2\lambda^3
$$

$$
= (2p - 1)\lambda - C'C(p)(2p - 1)^2\lambda^3 - (2p - 1)(\lambda - \lambda') + C'C(p)(2p - 1)^2(\lambda^3 - \lambda'^3)
$$

$$
\geq (2p - 1)\lambda - C'C(p)(2p - 1)^2\lambda^3 - \left[(2p - 1) + 3C'C(p)(2p - 1)^2\right] \epsilon,
$$

(5.7)

taking $\epsilon$ small enough we obtain the result for some constant $D(p) > 0$.

\[ \square \]

Remark 5.1. Notice that for a rational $\lambda$, by taking a uniform shift on the environment $\omega$ (shift right by an integer number uniformly chosen between 0 and the period of $\omega$), one gets an ergodic environment. Thus from Proposition 1.6, we get an example of a RWRE which achieves the speed bound up to $\lambda^3$.

5.2 Lack of tightness

We now present an example where no environment achieves the speed bound. This section also shows the bound in Proposition 1.6 can’t be improved asymptotically.

Let $p = 1$, $\lambda = \frac{n}{mn+1}$ and assume $1 < n \in \mathbb{N}$, $l \in \mathbb{N}$, $0 < l < n$ and $m \in \mathbb{N}$. Note that the assumptions hold for every rational number not of the form $\frac{1}{m}$ for some $m \in \mathbb{N}$, and that $\left\lceil \frac{1}{\lambda} \right\rceil = m + 1$, $\left\lfloor \frac{1}{\lambda} \right\rfloor = m$.

We prove the following proposition :

Proposition 5.2. For every environment $\nu \in \Upsilon$ the limit $\lim_{n \to \infty} X_n$ exists, and

$$
\lim_{n \to \infty} \frac{X_n}{n} = \frac{1}{2m + 1 - m(m + 1)\lambda}.
$$

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Proof. Since \( p = 1 \) we can write \( T_N \) as
\[
\frac{T_N}{N} = \frac{1}{N} \sum_{i=1}^{r_m(N)} S_i(m) + \frac{1}{N} \sum_{i=1}^{r_{m+1}(N)} S_i(m+1) + \frac{1}{N} S,
\]
where \( r_m(N), r_{m+1}(N) \) are the number of intervals of length \( m \) and \( m+1 \) up to time \( N \) respectively, \( \{S_i(m)\}_i \) and \( \{S_i(m+1)\}_i \) are two sequences of i.i.d random variables, where \( S_i(j) \) is distributed as the first hitting time of \( j \) by a simple random walk reflected at zero. \( S \) is the first hitting time to the point \( N - r_m(N)m - r_{m+1}(N)(m+1) \) of a simple random walk reflected at the origin, independently of both \( \{S_i(m)\}_i \) and \( \{S_i(m+1)\}_i \). Since \( S \) is finite almost surely and since \( \lim_{N \to \infty} \frac{r_i(N)}{N} = \rho_j \) for \( j \in \{m, m+1\} \) we get by the strong law of large numbers that
\[
\lim_{N \to \infty} \frac{T_N}{N} = \rho_m \cdot E[S_1(m)] + \rho_{m+1} E[S_1(m+1)]
\]
\[
= (m+1)\lambda - 1)m^2 + (1 - m\lambda)(m+1)^2
\]
\[
= 2m + 1 - m(m+1)\lambda.
\]
(5.9)

Following the same argument as in Lemma 2.1.17 in [8] we get that \( \lim_{N \to \infty} \frac{X_N}{N} \) exists and
\[
\lim_{N \to \infty} \frac{X_N}{N} = \frac{1}{\lim_{N \to \infty} \frac{T_N}{N}} = \frac{1}{2m + 1 - m(m+1)\lambda}.
\]

\[
\text{Proposition 5.3. For every } (1, \lambda) \text{ environment } \omega \text{ there exists an environment } \nu \in \Upsilon \text{ such that}
\]
\[
\liminf_{N \to \infty} \frac{T_N}{N} \geq \lim_{N \to \infty} \frac{T_N^\nu}{N},
\]
where \( T_N^\nu \) are the hitting times in the environment \( \nu \) and \( T_N \) are the ones in the environment \( \omega \).

Proof. Without loss of generality we assume that in the environment \( \omega \) the limit
\[
\lim_{n \to \infty} \frac{1}{n} \sum_{x=1}^{n} 1_{\omega(x) = p}
\]
exists and equals \( \lambda \). Indeed adding drifts to an environment only decreases the hitting times and we can always add drifts such that the limit of the density of the new environment exists. Let \( \{\epsilon_j\}_{j \in \mathbb{N}} \) a sequence of positive numbers such that \( \lim_{j \to \infty} \epsilon_j = 0 \). Notice that for large enough \( j \in \mathbb{N} \) we have \( m \equiv \left\lfloor \frac{1}{\lambda + \epsilon_j} \right\rfloor = \left\lfloor \frac{1}{\lambda} \right\rfloor \), and also \( m + 1 \equiv \left\lfloor \frac{1}{\lambda + \epsilon_j} \right\rfloor = \left\lceil \frac{1}{\lambda} \right\rceil \), and so we assume this is true for every \( j \in \mathbb{N} \). We turn now to define a sequence of environments \( \zeta^n \) which will gradually turn into an \( \Upsilon \) environment. We assume without loss of generality that \( \omega(0) = 1 \).

Fix some \( N_1 \in \mathbb{N} \) large enough such that for every \( n \geq N_1 - 1 \) the density in the interval \([0, n]\) is between \( \lambda - \epsilon_1 \) and \( \lambda + \epsilon_1 \). In particular we have
\[
\lambda - \epsilon_1 < \lambda_1 \equiv \frac{1}{N_1 - 1} \sum_{i=1}^{N_1-1} 1_{\omega(x) = p} < \lambda + \epsilon_1
\]
and also $\omega(N_1) = p$. Note that by the assumptions $\left\lfloor \frac{1}{\lambda_1} \right\rfloor = m$ and $\left\lceil \frac{1}{\lambda_1} \right\rceil = m + 1$.

Let us first analyze the environment in the interval $[0, N_1]$. We denote by $k_1$ the number of drifts in the interval $(0, N_1)$, and for $i \geq 1$ we denote by $r_i$ the number of intervals between two consecutive drifts of length $i$ (We count in the interval length the left drift but not the right one). We have the following relations:

$$N_1 = \sum_{i=0}^{N_1} ir_i$$
$$k_1 + 1 = \sum_{i=0}^{N_1} r_i$$
$$E[T_{N_1}] = \sum_{i=0}^{N_1} i^2 r_i. \quad (5.10)$$

Assume that there exist two indices $i < k$ such that $r_i, r_k > 0$, $k - i \geq 2$ and either $k > m + 1$ or $i < m$. By changing the location of the drifts one can replace one interval of length $k$ and one of length $i$ with intervals of length $k - 1$ and $i + 1$. By doing so one gets a new environment with the same total length, same number of drifts and with $E[T_N]$ smaller by $2(k - i - 1)$. Since the interval $[0, N_1]$ is finite, one can apply the last procedure only finite number of times, and achieve a new environment $\zeta^1$. Note that the environment $\zeta^1$ satisfy the following conditions:

- For every $x \geq N_1$ we have $\omega(x) = \zeta^1(x)$.
- In the environment $\zeta^1$ inside the interval $[0, N_1]$ there are only intervals of length $m$ and $m + 1$.

Indeed, the first claim is immediate from the fact the only changes we made where in the interval $(0, N_1)$. For the second claim, note that if in the end of the finite procedure one is left with an interval of length larger than $m + 1$, then all the intervals are of length larger or equal to $m + 1$ therefore the density is smaller than $\lambda_1$. Same argument shows no intervals of length smaller than $m$ are left at the end of the procedure in the interval $(0, N_1)$.

Since each step of the procedure defining $\zeta^1$ decreased the value of $E[T_{N_1}]$, and since $\omega$ and $\zeta^1$ coincide for $x \geq N_1$ we get that $E[T_n^{(1)}] \leq E[T_n]$ for every $n \geq N_1$, where $T_n^{(1)}$ is the first hitting time of where $n$ in the environment $\zeta^1$.

Let $N_2 \in \mathbb{N}$ be large enough so that $N_2 > N_1$,

$$\lambda - \epsilon_2 < \frac{1}{N_2 - 1} \sum_{i=1}^{N_2-1} 1_{\omega(x) = p} < \lambda + \epsilon_2,$$

and

$$\lambda - \epsilon_2 < \lambda_2 \equiv \frac{1}{N_2 - N_1 - 1} \sum_{i=N_1+1}^{N_2-1} 1_{\omega(x) = p} < \lambda + \epsilon_2.$$

Repeating the last procedure on the interval $[N_1, N_2]$ one can define a new environment $\zeta^2$ such that:
• $\zeta^2(x) = \omega(x)$ for every $x \geq N_2$.

• In the interval $[0, N_1]$ the environments $\zeta^1$ and $\zeta^2$ agree.

• In the interval $[0, N_2]$ the length between two consecutive drifts is either $m$ or $m + 1$.

• For every $n \geq N_2$ we have $E[T_n^{(2)}] \leq E[T_n^{(1)}] \leq E[T_n]$.

• For every $n \geq N_1$ the density of the drifts in the interval $(1, n)$ is between $\lambda - 2\epsilon_1$ and $\lambda + 2\epsilon_1$.

For the last point, notice that changing the order of intervals in $(N_1, N_2)$ does not change $E[T_n]$ for $n \geq N_2$. By rearranging the order of intervals we can ensure the last point is satisfied.

Repeating the last procedure and defining $\zeta^{j+1}$ from $\zeta^j$ in the same way, we get a sequence of environments. Finally define the environment $\nu$ by

$$
\nu(x) = \lim_{j \to \infty} \zeta^j(x), \quad \forall x \geq 0.
$$

This is well defined since for every $x \geq 0$ there exists $j_0 \in \mathbb{N}$ such that for every $j \geq j_0$ the value of $\zeta^j(x)$ is constant. From the definition of $\nu$ the environment is indeed in the family $\Upsilon$.

Denote by $l_i$ the location of the $i^{th}$ drift to the right of zero in the environment $\nu$ and $l_0 = 0$. In addition for every $n \in \mathbb{N}$ we define $k(n)$ to be the unique integer such that $l_{k(n)} < n \leq l_{k(n) + 1}$. It therefore follows that for every $n \in \mathbb{N}$ we have

$$
\frac{T_n}{n} = \frac{T_n - T_{l_{k(n)}}}{n} + \frac{1}{n} \sum_{i=1}^{k(n)} T_{l_i} - T_{l_{i-1}}.
$$

Since in the environment $\nu$ we only have intervals of length $m$ and $m + 1$ we have

$$
\frac{T_n - T_{l_{k(n)}}}{n} \leq \frac{T_n - T_{n-m-1}}{n}
$$

and therefore $\lim_{n \to \infty} \frac{T_n - T_{l_{k(n)}}}{n} = 0$, $\mathbb{P}_\omega$ a.s. Consequently we get that

$$
\liminf_{n \to \infty} \frac{T_n}{n} = \liminf_{n \to \infty} \frac{1}{n} \sum_{i=1}^{k(n)} T_{l_i} - T_{l_{i-1}}, \quad \mathbb{P}_\omega \text{ a.s.}
$$

Since $k(n)$ as defined above equals to the number of drifts in the interval $(0, n)$, we get from the construction of the environment $\nu$ that $\lim_{n \to \infty} \frac{k(n)}{n} = \lambda$. Thus we get that

$$
\liminf_{n \to \infty} \frac{T_n}{n} = \liminf_{n \to \infty} \frac{\lambda}{k(n)} \sum_{i=1}^{k(n)} T_{l_i} - T_{l_{i-1}}, \quad \mathbb{P}_\omega \text{ a.s.}
$$

which by Kolmogorov strong law of large numbers equals to

$$
\liminf_{n \to \infty} \frac{\lambda}{k(n)} \sum_{i=1}^{k(n)} \mathbb{E}\left[T_{l_i} - T_{l_{i-1}}\right] = \liminf_{n \to \infty} \frac{\lambda}{k(n)} \mathbb{E}\left[T_{l_{k(n)}}\right].
$$
Note that in order to apply Kolmogorov’s LLN we used the fact that \( l_i - l_{i-1} \leq \left\lceil \frac{1}{\lambda} \right\rceil \). Using again the construction of the environment \( \nu \) we get that the last expression is equal or bigger than

\[
\liminf_{n \to \infty} \frac{\lambda}{k(n)} E \left[ T_{l_k(n)}^\nu \right].
\]

Since in the environment \( \nu \), \( \lim_{n \to \infty} \frac{1}{n} \sum_{i=0}^{n-1} 1_{\nu(i)=1} \) exists and equals \( \lambda \) we have that \( \lim_{n \to \infty} \frac{l_k(n)}{k(n)} = \frac{1}{\lambda} \) and so we get that

\[
\liminf_{n \to \infty} \frac{T_n}{n} \geq \liminf_{n \to \infty} \frac{\lambda}{k(n)} E \left[ T_{l_k(n)}^\nu \right] = \liminf_{n \to \infty} \frac{1}{l_k(n)} E \left[ T_{l_k(n)}^\nu \right] \geq \liminf_{n \to \infty} \frac{E \left[ T_n^\nu \right]}{n}. \tag{5.11}
\]

By Proposition 5.2 the limit \( \lim_{n \to \infty} \frac{T_n}{n} = \frac{1}{\lim_{n \to \infty} X_n} \) exists, which together with Fatou lemma’s gives

\[
\liminf_{n \to \infty} \frac{E \left[ T_n^\nu \right]}{n} \geq E \left[ \liminf_{n \to \infty} \frac{T_n^\nu}{n} \right] = E \left[ \lim_{n \to \infty} \frac{T_n^\nu}{n} \right].
\]

Finally using Lemma 1.2 \( \lim_{n \to \infty} \frac{T_n^\nu}{n} \) is a \( \mathcal{P}_\nu \) almost sure constant. Thus

\[
E \left[ \lim_{n \to \infty} \frac{T_n^\nu}{n} \right] = \lim_{n \to \infty} \frac{T_n^\nu}{n}
\]

and

\[
\liminf_{n \to \infty} \frac{T_n}{n} \geq \lim_{n \to \infty} \frac{T_n^\nu}{n}, \quad \mathcal{P}_\omega,\nu \text{ a.s.}
\]

Proof of Proposition 1.7. By Proposition 5.3 there exists a \( \Upsilon \) environment \( \nu \) such that

\[
\liminf_{N \to \infty} \frac{T_N}{N} \geq \lim_{N \to \infty} \frac{T_N^\nu}{N},
\]

so it is enough to show that for every \( \Upsilon \) environment \( \nu \) we have

\[
\limsup_{N \to \infty} \frac{X_N}{N} \leq \lambda - D \lambda^3, \text{ a.s},
\]

for some constant \( D > 0 \). But this indeed holds since

\[
\lambda - \limsup_{N \to \infty} \frac{X_N}{N} = \lambda - \frac{1}{2m + 1 - m(m+1)} = \frac{n}{mn + l} - \frac{1}{2m + 1 - m(m+1)} \frac{n}{mn + l},
\]

rearranging the last expression we get

\[
= \lambda^3 \cdot \frac{l(n-l)}{n^2} \cdot \frac{1}{1 - l(n-l)}.
\]

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Using the fact that \( l > 0 \) and \( n > 1 \) we get that the last expression is bigger than

\[
\lambda^3 \cdot \left( \frac{1}{n} - \frac{1}{n^2} \right) \cdot \frac{1}{1 - \lambda^2 \left( \frac{1}{n} - \frac{1}{n^2} \right)} \geq \lambda^3 \cdot \left( \frac{1}{n} - \frac{1}{n^2} \right).
\]

\[\square\]

6 Transience Recurrence and the triviality of the \( \lim \sup \)

Definition 6.1. For a \((p, \lambda)\) environment \( \omega \) we define \( S(\omega) \) by

\[
S(\omega) = \sum_{n=1}^{\infty} \prod_{j=1}^{n} \rho(j),
\]

where as before for \( j \in \mathbb{N} \)

\[
\rho(j) = \frac{1 - \omega(j)}{\omega(j)}.
\]

Definition 6.2. For an environment \( \omega \) and \( x \in \mathbb{Z} \) we define \( \theta^x \omega \) to be the translation of \( \omega \) by \( x \), i.e. for every \( n \in \mathbb{Z} \), \( \theta^x \omega(n) = \omega(n + x) \).

Lemma 6.3. Fix a \((p, \lambda)\) environment \( \omega \). If \( S(\omega) < \infty \) then a random walk in \( \omega \) is transient to the right, i.e. for every \( x_0 \in \mathbb{Z} \) we have \( P_{\omega}^{x_0} (\lim_{n \to \infty} X_n = \infty) = 1 \). If \( S(\omega) = \infty \) then a random walk in \( \omega \) is recurrent, i.e. for every \( x_0 \in \mathbb{Z} \) we have \( P_{\omega}^{x_0} (-\infty = \liminf_{n \to \infty} X_n < \limsup_{n \to \infty} X_n = \infty) = 1 \).

Proof. This is a straight implication of the ideas and results of Theorem 2.1.2 of [8]. Note that since \( \omega(x) \geq \frac{1}{2} \) for all \( x \in \mathbb{Z} \), the walk can not be transient to the left. \[\square\]

Corollary 6.4. If for a \((p, \lambda)\) environment \( \omega \) the limit of the density exists and positive, i.e.

\[
\lim_{n \to \infty} \frac{1}{n} \sum_{i=0}^{n-1} \mathbb{1}_{\{\omega(i)=p\}} = \lambda > 0,
\]

then the random walk is transient to the right. Indeed, in this case one can fix \( 0 < \epsilon < \lambda \) and \( x_0 \in \mathbb{Z} \) and then find \( N \in \mathbb{N} \) such that for every \( n \geq N \) we have \( \frac{1}{n} \sum_{i=0}^{n-1} \mathbb{1}_{\{\omega(i)=p\}} > \lambda - \epsilon \). Thus for every \( n \geq N \)

\[
\sum_{k=1}^{n} \prod_{j=1}^{k} \rho(x+j) \leq \sum_{k=1}^{N-1} 1 + \sum_{k=N}^{n} \left( \frac{1-p}{p} \right)^{(\lambda-\epsilon)k} < N + \sum_{k=N}^{\infty} \left( \frac{1-p}{p} \right)^{(\lambda-\epsilon)k} < \infty
\]

since \( \frac{1-p}{p} < 1 \). Therefore by taking the limit \( n \to \infty \) one gets

\[
S(\theta^{x_0} \omega) < \infty
\]

and so the random walk is transient to the right.
Next we prove Lemma 1.2.

Proof of Lemma 1.2. For \( v \in \mathbb{R} \) and \( \delta > 0 \) we denote by \( A_{v,\delta} \) the event

\[
A_{v,\delta} = \left\{ \limsup_{n \to \infty} \frac{X_n}{n} - v < \delta \right\}.
\]

Assume that \( \mathbb{P}_\omega(A_{v,\delta}) > 0 \). Since \( A_{v,\delta} \in \sigma(X_1, X_2, \ldots) \), for every \( \epsilon > 0 \), one can find \( M \in \mathbb{N} \) and an event \( B_{v,\delta}^M \in \sigma(X_1, X_2, \ldots, X_M) \) such that

\[
\mathbb{P}_\omega(A_{v,\delta} \triangle B_{v,\delta}^M) < \epsilon.
\]

Notice that for small enough \( \epsilon > 0 \) this implies that \( \mathbb{P}_\omega(B_{v,\delta}^M) > \frac{\mathbb{P}_\omega(A_{v,\delta})}{2} \equiv c > 0 \).

Since \( X_n \) is a nearest neighbor random walk on \( \mathbb{Z} \) which starts at the origin we have the estimate \( |X_n| \leq n \), and therefore

\[
\begin{align*}
\mathbb{P}_\omega^0(A_{v,\delta} \cap B_{v,\delta}^M) &= \sum_{j=-M}^{M} \mathbb{P}_\omega^0(A_{v,\delta} \cap B_{v,\delta}^M \cap \{X_M = j\}) \\
&= \sum_{j=-M}^{M} \mathbb{P}_\omega^0(A_{v,\delta} | B_{v,\delta}^M \cap \{X_M = j\}) \cdot \mathbb{P}_\omega^0(B_{v,\delta}^M \cap \{X_M = j\}).
\end{align*}
\]

By the Markov property of the random walk this equals to

\[
\sum_{j=-M}^{M} \mathbb{P}_\omega^j(A_{v,\delta}) \cdot \mathbb{P}_\omega^0(B_{v,\delta}^M \cap \{X_M = j\}).
\]

Dividing the last formula by \( \mathbb{P}_\omega^0(B_{v,\delta}^M) \) we see that

\[
\mathbb{P}_\omega^0(A_{v,\delta} | B_{v,\delta}^M) = \sum_{j=-M}^{M} \mathbb{P}_\omega^j(A_{v,\delta}) \cdot \mathbb{P}_\omega^0(X_M = j | B_{v,\delta}^M).
\]

By the choice of \( B_{v,\delta}^M \) we get that

\[
\begin{align*}
\mathbb{P}_\omega^0(A_{v,\delta} | B_{v,\delta}^M) &= \frac{\mathbb{P}_\omega(A_{v,\delta} \cap B_{v,\delta}^M)}{\mathbb{P}_\omega^0(B_{v,\delta}^M)} = \frac{\mathbb{P}_\omega^0(B_{v,\delta}^M \setminus A_{v,\delta})}{\mathbb{P}_\omega^0(B_{v,\delta}^M)} \\
&= 1 - \frac{\mathbb{P}_\omega^0(B_{v,\delta}^M \setminus A_{v,\delta})}{\mathbb{P}_\omega^0(B_{v,\delta}^M)} \geq 1 - \frac{\epsilon}{c}.
\end{align*}
\]

In addition we have that

\[
\sum_{j=-M}^{M} \mathbb{P}_\omega^0(X_M = j | B_{v,\delta}^M) = 1.
\]
Using the last two observations and equation (6.2) we get that for small enough $\epsilon > 0$ there exists $M \in \mathbb{N}$ and $-M \leq j \leq M$ such that

$$P_j^{\omega}(A_{v,\delta}) > 1 - \frac{\epsilon}{c} > \frac{1}{2}.$$ 

Assume now towards contradiction that there exist two different values $v_1$ and $v_2$ in the support of $\limsup_{n \to \infty} X_n$. Choose $\delta_1, \delta_2 > 0$ small enough so that $A_{v_1,\delta_1} \cap A_{v_2,\delta_2} = \emptyset$. Using the conclusion of what we showed so far, one can find two integers $j_1$ and $j_2$ such that

$$P_{j_1}^{\omega}(A_{v_1,\delta_1}) > \frac{1}{2} \quad \text{and} \quad P_{j_2}^{\omega}(A_{v_2,\delta_2}) > \frac{1}{2}.$$ 

Without lost of generality we assume that $j_1 < j_2$. But according to Lemma 6.3 a random walk in a $(p, \lambda)$ environment $\omega$ is $P^{x}_\omega$ almost surely transient to the right or $P^{x}_\omega$ almost surely recurrent. and therefore a random walk starting at $j_1$ will reach $j_2$ at some finite random time $N$ almost surely. Consequently, if $X_n$ indeed starts at $j_1$, then

$$\limsup_{n \to \infty} \frac{X_n}{n} = \limsup_{n \to \infty} \frac{X_{n+N}}{n+N} = \limsup_{n \to \infty} \frac{X_{n+N}}{n}.$$ 

But the $\limsup$ on the left is distributed according to a random walk starting at $j_1$ and the one on the right is distributed according to a random walk starting at $j_2$, which gives the desired contradiction. \qed

7 Some conjectures and questions

In this article we studied random walks in $\mathbb{Z}$ environment composed of two point types, $(\frac{1}{2}, \frac{1}{2})$ and $(p, 1-p)$ for $p > \frac{1}{2}$. We ask for the following generalizations:

**Question 7.1.** What can be said about random walks in environments of $\mathbb{Z}$ composed of two types $(p, 1-p)$ and $(q, 1-q)$ for $\frac{1}{2} < p < q < 1$? More precisely we ask for a bound on the speed and give the following conjecture:

**Conjecture 7.2.** An environment which maximize the speed is given up to some integer effect by equally spaced drifts.

**Question 7.3.** What can be said about the speed of random walks with more than one type of drifts? For example about environments composed of three types $(\frac{1}{2}, \frac{1}{2})$, $(p, 1-p)$ and $(q, 1-q)$ for $\frac{1}{2} < p < q < 1$.

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References

[1] N. Berger, N. Kapur, L.J. Schulman, and V. Vazirani. Solvency games. In Proc. of FSTTCS, volume 8. Citeseer, 2008.

[2] Noam. Berger and Eviatar.B. Procaccia. Mutually excited random walk, in preperation.

[3] George T. Gilber. Positive definite matrices and sylvester’s criterion. Am. Math. Monthly, 98:44–46, January 1991.

[4] Gene H. Golub and Charles F. Van Loan. Matrix Computations. The Johns Hopkins University Press, 3rd edition, 1996.

[5] S. Lee. Optimal drift on (0, 1). Transactions of the American Mathematical Society, 346(1):159–176, 1994.

[6] M.V. Menshikov and A.R. Wade. Logarithmic speeds for one-dimensional perturbed random walks in random environments. Stochastic Processes and their Applications, 118(3):389–416, 2008.

[7] P. Morters and Y. Peres. Brownian motion, volume 30. Cambridge Univ Pr, 2010.

[8] O. Zeitouni. Lecture notes on random walks in random environment. St Flour Summer School, 2001.