Multidimensional Hungarian construction
for vectors with almost Gaussian smooth distributions

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Abstract: A multidimensional version of the results of Komlós, Major and Tusnády for sums of independent random vectors with finite exponential moments is obtained in the particular case where the summands have smooth distributions which are close to Gaussian ones. The bounds obtained reflect this closeness. Furthermore, the results provide sufficient conditions for the existence of i.i.d. vectors $X_1, X_2, \ldots$ with given distributions and corresponding i.i.d. Gaussian vectors $Y_1, Y_2, \ldots$ such that, for given small $\varepsilon$,

$$
P\left\{ \limsup_{n \to \infty} \frac{1}{\log n} \left| \sum_{j=1}^{n} X_j - \sum_{j=1}^{n} Y_j \right| \leq \varepsilon \right\} = 1.
$$

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1 Introduction

The paper is devoted to an improvement of a multidimensional version of strong approximation results of Komlós, Major and Tusnády (KMT) for sums of independent random vectors with finite exponential moments and with smooth distributions which are close to Gaussian ones.

Let $\mathcal{F}_d$ be the set of all $d$-dimensional probability distributions defined on the $\sigma$-algebra $\mathcal{B}_d$ of Borel subsets of $\mathbb{R}^d$. By $\hat{F}(t)$, $t \in \mathbb{R}^d$, we denote the characteristic function of a distribution $F \in \mathcal{F}_d$. The product of measures is understood as their convolution, that is, $FG = F \ast G$. The distribution and the corresponding covariance operator of a random vector $\xi$ will be denoted by $\mathcal{L}(\xi)$ and $\text{cov} \, \xi$ (or $\text{cov} \, F$, if $F = \mathcal{L}(\xi)$). The symbol $I_d$ will be used for the identity operator in $\mathbb{R}^d$. For $b > 0$ we denote $\log^* b = \max \{1, \log b\}$. Writing

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$z \in \mathbb{R}^d$ (resp. $\mathbb{C}^d$), we shall use the representation $z = (z_1, \ldots, z_d) = z_1 e_1 + \cdots + z_d e_d$, where $z_j \in \mathbb{R}^1$ (resp. $\mathbb{C}^1$) and the $e_j$ are the standard orthonormal vectors. The scalar product is denoted by $\langle x, y \rangle = x_1 y_1 + \cdots + x_d y_d$. We shall use the Euclidean norm $\|z\| = (z, z)^{1/2}$ and the maximum norm $|z| = \max_{1 \leq j \leq k} |z_j|$. The symbols $c, c_1, c_2, \ldots$ will be used for absolute positive constants. The letter $c$ may denote different constants when we do not need to fix their numerical values. The ends of proofs will be denoted by $\Box$.

Let us consider the definition and some useful properties of classes of distributions $A_d(\tau) \subset \mathcal{F}_d$, $\tau \geq 0$, introduced in Zaitsev (1986), see as well Zaitsev (1995, 1996, 1998a). The class $A_d(\tau)$ (with a fixed $\tau \geq 0$) consists of distributions $F \in \mathcal{F}_d$ for which the function

$$\varphi(z) = \varphi(F, z) = \log \int_{\mathbb{R}^d} e^{(z, x)} F \{dx\} \quad (\varphi(0) = 0)$$

is defined and analytic for $\|z\| \tau < 1$, $z \in \mathbb{C}^d$, and

$$|d_u d_v \varphi(z)| \leq \|u\| \tau \langle D v, v \rangle \quad \text{for all } u, v \in \mathbb{R}^d \text{ and } \|z\| \tau < 1,$$

where $D = \text{cov} F$, and the derivative $d_u \varphi$ is given by

$$d_u \varphi(z) = \lim_{\beta \to 0} \frac{\varphi(z + \beta u) - \varphi(z)}{\beta}.$$

It is easy to see that $\tau_1 < \tau_2$ implies $A_d(\tau_1) \subset A_d(\tau_2)$. Moreover, the class $A_d(\tau)$ is closed with respect to convolution: if $F_1, F_2 \in A_d(\tau)$, then $F_1 F_2 \in A_d(\tau)$. The class $A_d(0)$ coincides with the class of all Gaussian distributions in $\mathbb{R}^d$. The following inequality can be considered as an estimate of the stability of this characterization: if $F \in A_d(\tau)$, $\tau > 0$, then

$$\pi(F, \Phi(F)) \leq cd^2 \tau \log^*(\tau^{-1}), \quad (1.1)$$

where $\pi(\cdot, \cdot)$ is the Prokhorov distance and $\Phi(F)$ denotes the Gaussian distribution whose mean and covariance operator are the same as those of $F$. Moreover, for all $X \in \mathcal{B}_d$ and all $\lambda > 0$, we have

$$F\{X\} \leq \Phi(F)\{X^\lambda\} + cd^2 \exp \left( - \frac{\lambda}{cd^2 \tau} \right), \quad (1.2)$$

$$\Phi(F)\{X\} \leq F\{X^\lambda\} + cd^2 \exp \left( - \frac{\lambda}{cd^2 \tau} \right), \quad (1.3)$$

where $X^\lambda = \{ y \in \mathbb{R}^d : \inf_{x \in X} \|x - y\| < \lambda \}$ is the $\lambda$-neighborhood of the set $X$, see Zaitsev (1986).

The classes $A_d(\tau)$ are closely connected with other natural classes of multidimensional distributions. In particular, by the definition of $A_d(\tau)$, any distribution $\mathcal{L}(\xi)$ from $A_d(\tau)$ has finite exponential moments $\mathbb{E} e^{(h, \xi)}$, for $\|h\| \tau < 1$. This leads to exponential estimates for the tails of distributions (see, e.g., Lemma 3.3 below). On the other hand, if $\mathbb{E} e^{(h, \xi)} < \infty$, for $h \in A \subset \mathbb{R}^d$, where $A$ is a neighborhood of zero, then $F = \mathcal{L}(\xi) \in A_d(\tau(F))$ with some $\tau(F)$ depending on $F$ only.
Throughout we assume that \( \tau \geq 0 \) and \( \xi_1, \xi_2, \ldots \) are random vectors with given distributions \( L(\xi_k) \in A_d(\tau) \) such that \( E\xi_k = 0 \), \( \text{cov}\xi_k = I_d \), \( k = 1, 2, \ldots \). The problem is to construct, for a given \( n \), \( 1 \leq n \leq \infty \), on a probability space a sequence of independent random vectors \( X_1, \ldots, X_n \) and a sequence of i.i.d. Gaussian random vectors \( Y_1, \ldots, Y_n \) with \( L(X_k) = L(\xi_k), \ EY_k = 0, \text{cov}Y_k = I_d, \ k = 1, \ldots, n \), such that, with large probability,

\[
\Delta(n) = \max_{1 \leq r \leq n} \left| \sum_{k=1}^{r} X_k - \sum_{k=1}^{r} Y_k \right|
\]

is as small as possible.

The aim of the paper is to provide sufficient conditions for the following Assertion A:

**Assertion A.** There exist absolute positive constants \( c_1, c_2 \) and \( c_3 \) such that, for \( \tau d^{3/2} \leq c_1 \), there exists a construction with

\[
E \exp \left( \frac{c_2 \Delta(n)}{d^{3/2} \tau} \right) \leq \exp \left( c_3 \log^* d \log^* n \right).
\]  

(1.4)

Using the exponential Chebyshev inequality, we see that (1.4) implies

\[
P \left\{ c_2 \Delta(n) \geq \tau d^{3/2} \left( c_3 \log^* d \log^* n + x \right) \right\} \leq e^{-x}, \quad x \geq 0.
\]

(1.5)

Therefore, Assertion A can be considered as a generalization of the classical result of KMT (1975, 1976). Assertion A provides a supplement to an improvement of a multidimensional KMT-type result of Einmahl (1989) presented by Zaitsev (1995, 1998a) which differs from Assertion A by the restriction \( \tau \geq 1 \) and by another explicit power-type dependence of the constants on the dimension \( d \). In a particular case, when \( d = 1 \) and all summands have a common variance, the result of Zaitsev is equivalent to the main result of Sakhanenko (1984), who extended the KMT construction to the case of non-identically distributed summands and stated the dependence of constants on the distributions of the summands belonging to a subclass of \( A_1(\tau) \). The main difference between Assertion A and the aforementioned results consists in the fact that in Assertion A we consider "small" \( \tau \), \( 0 \leq \tau \leq c_1 d^{-3/2} \).

In previous results the constants are separated from zero by quantities which are larger than some absolute constants. In KMT (1975, 1976) the dependence of the constants on the distributions is not specified. From the conditions (1) and (4) in Sakhanenko (1984, Section 1), it follows that \( \text{Var}\xi_k \leq \lambda^{-2} \) (\( \lambda^{-1} \) plays in Sakhanenko’s paper the role of \( \tau \)) and, if \( \text{Var}\xi_k = 1 \), then \( \lambda^{-1} \geq 1 \). This corresponds to the restrictions \( \alpha^{-1} \geq 2 \) in Einmahl (1989, conditions (3.6) and (4.3)) and \( \tau \geq 1 \) in Zaitsev (1995, 1998a, Theorem 1).

Note that in Assertion A we do not require that the distributions \( L(\xi_k) \) are identical but we assume that they have the same covariance operators, cf. Einmahl (1989) and Zaitsev (1995, 1998a). A generalization of the results of Zaitsev (1995, 1998a) and of the present paper to the case of non-identical covariance operators appeared recently in the preprint Zaitsev (1998b).

According to (1.1)–(1.3), the condition \( L(\xi_k) \in A_d(\tau) \) with small \( \tau \) means that \( L(\xi_k) \) are close to the corresponding Gaussian laws. It is easy to see that Assertion A becomes stronger for small \( \tau \) (see as well Theorem 1.4 below). Passing to the limit as \( \tau \to 0 \), we
obtain a spectrum of statements with the trivial limiting case: if \( \tau = 0 \) (and, hence, \( \mathcal{L}(\xi_k) \) are Gaussian) we can take \( X_k = Y_k \) and \( \Delta(n) = 0 \).

We show that Assertion A is valid under some additional smoothness-type restrictions on \( \mathcal{L}(\xi_k) \). The question about the necessity of these conditions remains open. The case \( \tau \geq 1 \) considered by Zaitsev (1995, 1998a, Theorem 1) does not need conditions of such kind. The formulation of our main result—Theorem 2.1—includes some additional notation.

In order to show that the conditions of Theorem 2.1 can be verified in some concrete simple situations, we consider at first three particular applications—Theorems 1.1, 1.2 and 1.3.

**Theorem 1.1.** Assume that the distributions \( \mathcal{L}(\xi_k) \in \mathcal{A}_d(\tau) \) can be represented in the form

\[
\mathcal{L}(\xi_k) = H_k G, \quad k = 1, \ldots, n,
\]

where \( G \) is a Gaussian distribution with covariance operator \( \text{cov} G = b^2 I_d \) with \( b^2 \geq 2^{10} \tau^2 d^3 \log^* \frac{1}{\tau} \). Then Assertion A is valid.

The following example deals with a non-convolution family of distributions approximating a Gaussian distribution for small \( \tau \).

**Theorem 1.2.** Let \( \eta \) be a random vector with an absolutely continuous distribution and density

\[
p_\tau(x) = \frac{(4 + \tau^2 \|x\|^2) \exp \left( \frac{-\|x\|^2}{2} \right)}{(2\pi)^{d/2}(4 + \tau^2 d)}, \quad x \in \mathbb{R}^d.
\]

(1.6)

Assume that \( \mathcal{L}(\xi_k) = \mathcal{L}(\eta/\gamma), \quad k = 1, \ldots, n \), where

\[
\gamma^2 = \frac{(4 + \tau^2 (d + 2))}{(4 + \tau^2 d)}, \quad \gamma > 0.
\]

(1.7)

Then Assertion A is valid.

The proof of Theorem 1.2 can be apparently extended to the distributions with some more general densities of type \( P(\tau^2 \|x\|^2) \exp \left( -c \|x\|^2 \right) \), where \( P(\cdot) \) is a suitable polynomial.

**Theorem 1.3.** Assume that a random vector \( \zeta \) satisfies the relations

\[
\mathbb{E} \zeta = 0, \quad \mathbb{P} \left\{ \| \zeta \| \leq b_1 \right\} = 1, \quad H := \mathcal{L}(\zeta) \in \mathcal{A}_d(b_2)
\]

(1.8)

and admits a differentiable density \( p(\cdot) \) such that

\[
\sup_{x \in \mathbb{R}^d} \left| d_u p(x) \right| \leq b_3 \| u \|, \quad \text{for all } u \in \mathbb{R}^d,
\]

(1.9)

with some positive \( b_1, b_2 \) and \( b_3 \). Let \( \zeta_1, \zeta_2, \ldots \) be independent copies of \( \zeta \). Write

\[
\tau = b_2 m^{-1/2},
\]

(1.10)

where \( m \) is a positive integer. Assume that the distributions \( \mathcal{L}(\xi_k) \) can be represented in the form

\[
\mathcal{L}(\xi_k) = L^{(k)} P, \quad k = 1, \ldots, n,
\]

(1.11)

where

\[
L^{(k)} \in \mathcal{A}_d(\tau) \quad \text{and} \quad P = \mathcal{L} \left( \left( \zeta_1 + \cdots + \zeta_m \right)/\sqrt{m} \right).
\]

(1.12)

Then there exist a positive \( b_4 \) depending on \( H \) only and such that \( m \geq b_4 \) implies Assertion A.
Remark 1.1. If all the distributions \( L^{(k)} \) are concentrated at zero, then the statement of Theorem 1.3 (for \( \tau = bm^{-1/2} \) with some \( b = b(H) \)) can be derived from the main results of KMT (1975, 1976) (for \( d = 1 \)) and of Zaitsev (1995, 1998a) (for \( d \geq 1 \)).

A consequence of Assertion A is given in Theorem 1.4 below.

**Theorem 1.4.** Assume that \( \xi, \xi_1, \xi_2, \ldots \) are i.i.d. random vectors with a common distribution \( L(\xi) \in A_d(\tau) \). Let Assertion A be satisfied for \( \xi_1, \ldots, \xi_n \) for all \( n \) with some \( c_1, c_2 \) and \( c_3 \) independent of \( n \). Suppose that \( \tau d^{3/2} \leq c_1 \). Then there exist a construction such that
\[
P\left\{ \limsup_{n \to \infty} \frac{1}{\log n} \left| \sum_{j=1}^{n} X_j - \sum_{j=1}^{n} Y_j \right| \leq c_4 \tau d^{3/2} \log^* d \right\} = 1 \tag{1.13}
\]
with some constant \( c_4 = c_4(c_2, c_3) \).

From a result of Bártfai (1966) it follows that the rate \( O(\log n) \) in (1.13) is the best possible if \( L(\xi) \) is non-Gaussian. In the case of distributions with finite exponential moments this rate was established by Zaitsev (1995, 1998a, Corollary 1). Theorems 1.1–1.3 and 2.1 provide examples of smooth distributions which are close to Gaussian ones and for which the constants corresponding to this rate are arbitrarily small. The existence of such examples has been already mentioned in the one-dimensional case, e.g., by Major (1978, p. 498).

The paper is organized as follows. In Section 2 we formulate Theorem 2.1. To this end we define at first a class of distributions \( \mathcal{A}_d(\tau, \rho) \) used in Theorem 2.1. The definition of this class is given in terms of smoothness conditions on the so-called conjugate distributions. Then we describe a multidimensional version of the KMT dyadic scheme, cf. Einmahl (1989). We prove Theorem 2.1 in Section 3. Section 4 is devoted to the proofs of Theorems 1.1–1.4.

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## 2 The main result

Let \( F = L(\xi) \in \mathcal{A}_d(\tau) \), \( \|h\| \tau < 1 \), \( h \in \mathbb{R}^d \). The conjugate distribution \( \overline{F} = \overline{F}(h) \) is defined by
\[
\overline{F}\{dx\} = (\mathbb{E} e^{\langle h, \xi \rangle})^{-1} e^{\langle h, x \rangle} F\{dx\}. \tag{2.1}
\]
Sometimes we shall write \( F_h = \overline{F}(h) \). It is clear that \( F(0) = F \). Denote by \( \overline{X}(h) \) a random vector with \( L(\overline{X}(h)) = \overline{F}(h) \). From (2.1) it follows that
\[
\mathbb{E} f(\overline{X}(h)) = (\mathbb{E} e^{\langle h, \xi \rangle})^{-1} \mathbb{E} f(\xi) e^{\langle h, \xi \rangle}, \tag{2.2}
\]
provided that \( \mathbb{E} |f(\xi) e^{\langle h, \xi \rangle}| < \infty \). It is easy to see that
\[
\text{if } U_1, U_2 \in \mathcal{A}_d(\tau), \ U = U_1 U_2, \ \text{then } \overline{U}(h) = \overline{U_1}(h) \overline{U_2}(h). \tag{2.3}
\]
In particular, let a collection of summands with indices of the form $0 \leq k < 2^n$. Consider the conditions:

\[ \int_{\rho\|t\| \tau d \geq 1} \left| \hat{F}_h(t) \right| dt \leq \frac{(2\pi)^{d/2} \tau^3}{\sigma (\text{det} D)^{1/2}}, \quad (2.4) \]

\[ \int_{\rho\|t\| \tau d \geq 1} \left| \hat{F}_h(t) \right| dt \leq \frac{(2\pi)^{d/2} \tau^2 d^2}{\sigma^2 (\text{det} D)^{1/2}}, \quad (2.5) \]

\[ \int_{\rho\|t\| \tau d \geq 1} \left| \langle t, v \rangle \hat{F}_h(t) \right| dt \leq \frac{(2\pi)^{d/2} \langle D^{-1} v, v \rangle^{1/2}}{\delta (\text{det} D)^{1/2}}, \quad \text{for all } v \in \mathbb{R}^d, \quad (2.6) \]

where $F_h = \overline{F}(h)$ and $\sigma^2 = \sigma^2(F) > 0$ is the minimal eigenvalue of $D = \text{cov} F$. Denote by $\mathcal{A}_d(\tau, \rho)$ (resp. $\mathcal{A}_d^*(\tau, \delta, \rho)$) the class of distributions $F \in \mathcal{A}_d(\tau)$ such that the condition (2.4) (resp. (2.5) and (2.6)) is satisfied for $h \in \mathbb{R}^d, \|h\| \tau < 1$. It is easy to see that

\[ \mathcal{A}_d^*(\tau, \delta, \rho) \subset \overline{\mathcal{A}_d(\tau, \rho)}, \quad \text{if } \frac{\tau d^{1/2}}{\sigma} \leq 1. \quad (2.7) \]

In this paper the class $\overline{\mathcal{A}_d(\tau, \rho)}$ plays the role of the class $\mathcal{A}_d^*(\tau, \delta, \rho)$ which was used by Zaitsev (1995, 1998a), see also Sakhanenko (1984, inequality (49), p. 9) or Einmahl (1989, inequality (1.5)). Note that (2.2) implies

\[ \hat{F}_h(t) = E e^{i(t, \xi(h))} = \left( E e^{(h, \xi)} \right)^{-1} E e^{(h+it, \xi)}. \quad (2.8) \]

**The dyadic scheme.** Let $N$ be a positive integer and $\{ \xi_1, \ldots, \xi_{2^N} \}$ a collection of $d$-dimensional independent random vectors. Denote

\[ \tilde{S}_0 = 0; \quad \tilde{S}_k = \sum_{i=1}^{k} \xi_i, \quad 1 \leq k \leq 2^N; \quad (2.9) \]

\[ U_{m,k}^* = \tilde{S}_{(k+1) \cdot 2^m} - \tilde{S}_{k \cdot 2^m}, \quad 0 \leq k < 2^{N-m}, \quad 0 \leq m \leq N. \quad (2.10) \]

In particular, $U_{0,k}^* = \xi_{k+1}, U_{N,0}^* = \tilde{S}_{2^N} = \xi_1 + \cdots + \xi_{2^N}$. In the sequel we call block of summands a collection of summands with indices of the form $k \cdot 2^m + 1, \ldots, (k+1) \cdot 2^m$, where $0 \leq k < 2^{N-m}, \quad 0 \leq m \leq N$. Thus, $U_{m,k}^*$ is the sum over a block containing $2^m$ summands. Put

\[ \tilde{U}_{n,k}^* = U_{n-1,2k}^* - U_{n-1,2k+1}^*, \quad 0 \leq k < 2^{N-n}, \quad 1 \leq n \leq N. \quad (2.11) \]

Note that

\[ U_{n-1,2k}^* + U_{n-1,2k+1}^* = U_{n,k}^*, \quad 0 \leq k < 2^{N-n}, \quad 1 \leq n \leq N. \quad (2.12) \]
Introduce the vectors
\[ \tilde{U}_{n,k}^* = \left( U_{n-1,2,k}^*, U_{n-1,2,k+1}^* \right) \in \mathbb{R}^{2d}, \quad 0 \leq k < 2^{N-n}, \quad 1 \leq n \leq N, \] (2.13)
with the first \( d \) coordinates coinciding with those of the vectors \( U_{n-1,2,k}^* \) and with the last \( d \) coordinates coinciding with those of the vectors \( U_{n-1,2,k+1}^* \). Similarly, denote
\[ U_{n,k}^* = \left( U_{n,k}^*, \tilde{U}_{n,k}^* \right) \in \mathbb{R}^{2d}, \quad 0 \leq k < 2^{N-n}, \quad 1 \leq n \leq N. \] (2.14)

Introduce now the projectors \( P_i : \mathbb{R}^s \to \mathbb{R}^1 \) and \( \overline{P}_j : \mathbb{R}^s \to \mathbb{R}^j \), for \( i, j = 1, \ldots, s \), by the relations \( P_i x = x_i, \overline{P}_j x = (x_1, \ldots, x_j) \), where \( x = (x_1, \ldots, x_s) \in \mathbb{R}^s \) (we shall use this notation for \( s = d \) or \( s = 2d \)).

It is easy to see that, according to (2.11)–(2.14),
\[ U_{n,k}^* = A \tilde{U}_{n,k}^* \in \mathbb{R}^{2d}, \quad 0 \leq k < 2^{N-n}, \quad 1 \leq n \leq N, \] (2.15)
where \( A : \mathbb{R}^{2d} \to \mathbb{R}^{2d} \) is a linear operator defined, for \( x = (x_1, \ldots, x_{2d}) \in \mathbb{R}^{2d} \), as follows:
\[ P_j A x = x_j + x_{d+j}, \quad j = 1, \ldots, d, \]
\[ \overline{P}_j A x = x_j - x_{d+j}, \quad j = d + 1, \ldots, 2d. \] (2.16)

Denote
\[ U_{n,k}^{*j} = P_j U_{n,k}^*, \quad U_{n,k}^{*j} = \left( U_{n,k}^{*(1)}, \ldots, U_{n,k}^{*(j)} \right) = \overline{P}_j U_{n,k}^* \in \mathbb{R}^j, \quad j = 1, \ldots, 2d. \] (2.17)

Now we can formulate the main result of the paper.

**Theorem 2.1.** Let the conditions described in (2.9)–(2.17) be satisfied, \( \tau \geq 0 \) and \( E \xi_k = 0 \), \( \text{cov} \xi_k = I_d, \) \( k = 1, \ldots, 2^N \). Assume that
\[ \mathcal{L}(U_{n,k}^{*j}) \in \overline{A}_j(\tau, 4) \quad \text{for} \quad 0 \leq k < 2^{N-n}, \quad 1 \leq n \leq N, \quad d \leq j \leq 2d, \] (2.18)
and
\[ \mathcal{L}(U_{n,0}^{*j}) \in \overline{A}_j(\tau, 4) \quad \text{for} \quad 1 \leq j \leq 2d. \] (2.19)

Then there exist absolute positive constants \( c_5, c_6 \) and \( c_7 \) such that, for \( \tau d^{3/2} \leq c_5 \), one can construct on a probability space sequences of independent random vectors \( X_1, \ldots, X_{2^N} \) and i.i.d. Gaussian random vectors \( Y_1, \ldots, Y_{2^N} \) so that
\[ \mathcal{L}(X_k) = \mathcal{L}(\xi_k), \quad E Y_k = 0, \quad \text{cov} Y_k = I_d, \quad k = 1, \ldots, 2^N, \] (2.20)
and
\[ E \exp\left( \frac{c_6 \Delta(2^N)}{d^{3/2}\tau} \right) \leq \exp \left( c_7 N \log^* d \right), \] (2.21)
where \( \Delta(2^N) = \max_{1 \leq r \leq 2^N} \left| \sum_{k=1}^r X_k - \sum_{k=1}^r Y_k \right|. \)
Theorem 2.1 says that the conditions (2.18) and (2.19) suffice for Assertion A. However, these conditions require that the number of summands is \( 2^N \). For an arbitrary number of summands, one should consider additional (for simplicity, Gaussian) summands in order to apply Theorem 2.1.

Below we shall prove Theorem 2.1. Suppose that its conditions are satisfied.

At first, we describe a procedure of constructing the random vectors \( \{ U_{n,k} \} \) with distributions \( \mathcal{L}\{ U_{n,k} \} = \mathcal{L}\{ U^*_{n,k} \} \), provided that the vectors \( Y_1, \ldots, Y_{2^N} \) are already constructed (then we shall define \( X_k = U_{0,k-1}, \ k = 1, \ldots, 2^N \)). This procedure is an extension of the KMT (1975, 1976) dyadic scheme to the multivariate case due to Einmahl (1989). For this purpose we shall use the so-called Rosenblatt quantile transformation (see Rosenblatt (1952) and Einmahl (1989)).

Denote by \( F_{N,0}^{(1)}(x_1) = P\{ P_1 U^*_{N,0} < x_1 \} \), \( x_1 \in \mathbb{R}^1 \), the distribution function of the first coordinate of the vector \( U^*_{N,0} \). Introduce the conditional distributions, denoting by \( F_{N,0}^{(j)}( \cdot \mid x_1, \ldots, x_{j-1} ) \), \( 2 \leq j \leq d \), the regular conditional distribution function (r.c.d.f.) of \( P_j U^*_{N,0} \), given \( P_{j-1} U^*_{N,0} = (x_1, \ldots, x_{j-1}) \). Denote by \( \bar{F}_{n,k}^{(j)}( \cdot \mid x_1, \ldots, x_{j-1} ) \) the r.c.d.f. of \( P_j U^*_{n,k} \), given \( P_{j-1} U^*_{n,k} = (x_1, \ldots, x_{j-1}) \), for \( 0 \leq k < 2^{N-n}, \ 1 \leq n \leq N, \ d+1 \leq j \leq 2d \). Put

\[
T_k = \sum_{l=1}^{k} Y_l, \quad 1 \leq k \leq 2^N; \tag{2.22}
\]

\[
V_{m,k} = (V^{(1)}_{m,k}, \ldots, V^{(d)}_{m,k}) = T_{(k+1)2^m} - T_{k2^m}, \quad 0 \leq k < 2^{N-m}, \ 0 \leq m \leq N; \tag{2.23}
\]

\[
\bar{V}_{n,k} = (V_{n-1,2k}, V_{n-1,2k+1}) = \left( \bar{V}^{(1)}_{n,k}, \ldots, \bar{V}^{(2d)}_{n,k} \right) \in \mathbb{R}^{2d},
\]

\[
0 \leq k < 2^{N-n}, \ 1 \leq n \leq N;
\]

and

\[
V_{n,k} = (V^{(1)}_{n,k}, \ldots, V^{(2d)}_{n,k}) = A \bar{V}_{n,k} \in \mathbb{R}^{2d}, \quad 0 \leq k < 2^{N-n}, \ 1 \leq n \leq N. \tag{2.24}
\]

According to the definition of the operator \( A \), we have (see (2.11)–(2.16) and (2.22)–(2.24))

\[
V_{n,k} = (V_{n,k}, \bar{V}_{n,k}) \in \mathbb{R}^{2d}, \quad 0 \leq k < 2^{N-n}, \ 1 \leq n \leq N, \tag{2.25}
\]

where

\[
V_{n,k} = V_{n-1,2k} + V_{n-1,2k+1}, \quad 0 \leq k < 2^{N-n}, \ 1 \leq n \leq N, \tag{2.26}
\]

and

\[
V_{N,0} = Y_1 + \cdots + Y_{2^N}. \tag{2.27}
\]

Thus, the vectors \( V_{m,k}, \bar{V}_{n,k} \) and \( V_{n,k} \) can be constructed from the vectors \( Y_1, \ldots, Y_{2^N} \) by the same linear procedure which was used for constructing the vectors \( U^*_{m,k}, \bar{U}^*_{n,k} \) and \( U^*_{n,k} \) from the vectors \( \xi_1, \ldots, \xi_{2^N} \).
It is obvious that, for fixed \( n \) and \( k \),
\[
\text{cov } U_{n,k}^* = \text{cov } V_{n,k} = 2^n I_{2d}
\]  
(2.28)
and, hence, the coordinates of the Gaussian vector \( V_{n,k} \) are independent with the same distribution function \( \Phi_{2^n/2}(\cdot) \) (here and below
\[
\Phi_\sigma(x) = \int_{-\infty}^{x} \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{y^2}{2\sigma^2}\right) \, dy, \quad x \in \mathbb{R}^1, \quad \sigma > 0,
\]
is the distribution function of the normal law with mean zero and variance \( \sigma^2 \).

Denote now the new collection of random vectors \( X_k \) as follows. At first we define
\[
U_{N,0}^{(1)} = (F_{N,0}^{(1)})^{-1}\left( \Phi_{2^{N/2}}(V_{N,0}^{(1)}) \right) \quad \text{and, for } 2 \leq j \leq d,
\]
(2.29)
\[
U_{N,0}^{(j)} = (F_{N,0}^{(j)})^{-1}\left( \Phi_{2^{N/2}}(V_{N,0}^{(j)}) \mid U_{N,0}^{(1)}, \ldots, U_{N,0}^{(j-1)} \right)
\]
(here \((F_{N,0}^{(1)})^{-1}(t) = \sup \{ x : F_{N,0}^{(1)}(x) \leq t \}, \ 0 < t < 1, \ \text{and so on}).Taking into account that the distributions of the random vectors \( \xi_1, \ldots, \xi_{2^N} \) are absolutely continuous, we see that this formula can be rewritten in a more natural form, cf. Sakhanenko (1984, p. 30–31):
\[
F_{N,0}^{(1)}(U_{N,0}^{(1)}) = \Phi_{2^{N/2}}(V_{N,0}^{(1)}), \quad F_{N,0}^{(j)}(U_{N,0}^{(j)} \mid U_{N,0}^{(1)}, \ldots, U_{N,0}^{(j-1)}) = \Phi_{2^{N/2}}(V_{N,0}^{(j)}), \quad \text{for } 2 \leq j \leq d.
\]
(2.30)
Suppose that the random vectors
\[
U_{n,k} = (U_{n,k}^{(1)}, \ldots, U_{n,k}^{(d)}), \quad 0 \leq k < 2^{N-n},
\]
(2.31)
corresponding to blocks containing each \( 2^n \) summands with fixed \( n, \ 1 \leq n \leq N \), are already constructed. Now our aim is to construct the blocks containing each \( 2^{n-1} \) summands. To this end we define
\[
U_{n,k}^{(j)} = P_j U_{n,k} = U_{n,k}^{(j)}, \quad 1 \leq j \leq d,
\]
(2.32)
and, for \( d+1 \leq j \leq 2d, \)
\[
U_{n,k}^{(j)} = (\tilde{F}_{n,k}^{(j)})^{-1}\left( \Phi_{2^{2n/2}}(V_{n,k}^{(j)}) \mid U_{n,k}^{(1)}, \ldots, U_{n,k}^{(j-1)} \right).
\]
(2.33)
It is clear that (2.33) can be rewritten in a form similar to (2.30). Then we put
\[
U_{n,k} = (U_{n,k}^{(1)}, \ldots, U_{n,k}^{(2d)}) \in \mathbb{R}^{2d},
\]
\[
U_{n,k}^{(j)} = (U_{n,k}^{(1)}, \ldots, U_{n,k}^{(j)}) = P_j U_{n,k} \in \mathbb{R}^j, \quad j = 1, \ldots, 2d,
\]
(2.34)
\[
\tilde{U}_{n,k}^{(j)} = U_{n,k}^{(j+d)}, \quad j = 1, \ldots, d,
\]
\[
\tilde{U}_{n,k} = (\tilde{U}_{n,k}^{(1)}, \ldots, \tilde{U}_{n,k}^{(d)}) \in \mathbb{R}^d
\]
ard
\[
U_{n-1,2k} = \left( U_{n,k} + \tilde{U}_{n,k} \right) / 2,
\]
\[
U_{n-1,2k+1} = \left( U_{n,k} - \tilde{U}_{n,k} \right) / 2.
\]
(2.35)
Thus, we have constructed the random vectors $U_{n-1,k}$, $0 \leq k < 2^{N-n+1}$. After $N$ steps we obtain the random vectors $U_{0,k}$, $0 \leq k < 2^N$. Now we set

$$X_k = U_{0,k-1}, \quad S_0 = 0, \quad S_k = \sum_{l=1}^{k} X_l, \quad 1 \leq k \leq 2^N. \quad (2.36)$$

**Lemma 2.1.** (Einmahl (1989)) The joint distribution of the vectors $U_{n,k}$ and $U_{n,k}$ coincides with that of the vectors $U_{n,k}^*$ and $U_{n,k}^*$. In particular, $X_k$, $k = 1, \ldots, 2^N$, are independent and $\mathcal{L}(X_k) = \mathcal{L}(\xi_k)$.

Moreover, according to (2.11) and (2.12), we have

$$\tilde{U}_{n,k} = U_{n-1,2k} - U_{n-1,2k+1}, \quad U_{n,k} = U_{n-1,2k} + U_{n-1,2k+1} = S_{(k+1)2^n} - S_{k2^n}, \quad (2.37)$$

for $0 \leq k < 2^{N-n}$, $1 \leq n \leq N$ (it is clear that (2.37) follows from (2.35)). Furthermore, putting

$$\tilde{U}_{n,k} = \left( U_{n-1,2k}, U_{n-1,2k+1} \right) \in \mathbb{R}^{2d}, \quad 0 \leq k < 2^{N-n}, \quad 1 \leq n \leq N, \quad (2.38)$$

we have (see (2.13) and (2.15))

$$U_{n,k} = A \tilde{U}_{n,k} \in \mathbb{R}^{2d}, \quad 0 \leq k < 2^{N-n}, \quad 1 \leq n \leq N. \quad (2.39)$$

Note that it is not difficult to verify that, according to (2.16),

$$\|A\| = \frac{1}{\|A^{-1}\|} = \|A^*\| = \frac{1}{\|(A^*)^{-1}\|} = \sqrt{2}, \quad (2.40)$$

where the asterisk is used to denote the adjoint operator $A^*$ for the operator $A$.

**Remark 2.1.** The conditions of Theorem 2.1 imply the coincidence of the corresponding first and second moments of the vectors $\mathbf{U} = \{U_{n,k}, \tilde{U}_{n,k}, U_{n,k}\}$ and $\mathbf{V} = \{V_{n,k}, \tilde{V}_{n,k}, V_{n,k}\}$ since the vectors $\mathbf{U}$ can be restored from vectors $X_1, \ldots, X_{2^N}$ by the same linear procedure which is used for reconstruction of the vectors $\mathbf{V}$ from $Y_1, \ldots, Y_{2^N}$. In particular, $\mathbf{E} \mathbf{U} = \mathbf{E} \mathbf{V} = 0$.

**Lemma 2.2.** (Einmahl 1989, Lemma 5, p. 55) Let $1 \leq m = (2s + 1) \cdot 2^r \leq 2^N$, where $s, r$ are non-negative integers. Then

$$S_m = \frac{m}{2^N} S_{2^N} + \sum_{n=r+1}^{N} \gamma_n \tilde{U}_{n,l_n,m}, \quad (2.41)$$

where $\gamma_n = \gamma_n(m) \in [0, \frac{1}{2}]$ and the integers $l_n, m$ are defined by

$$l_n \cdot 2^n < m \leq (l_n + 1) \cdot 2^n. \quad (2.42)$$
The shortest proof of Lemma 2.2 can be obtained with the help of a geometrical approach due to Massart (1989, p. 275).

**Remark 2.2.** The inequalities (2.42) give a formal definition of \( l_{n,m} \). To understand better the mechanism of the dyadic scheme, one should remember another characterization of these numbers: \( U_{n,l_{n,m}} \) is the sum over the block of \( 2^n \) summands which contains \( X_m \), the last summand in the sum \( S_m \).

**Corollary 2.1.** Under the conditions of Lemma 2.2

\[
|S_m - T_m| \leq \left| U_{N,0} - V_{N,0} \right| + \frac{1}{2} \sum_{n=r+1}^{N} \left| \tilde{U}_{n,l_{n,m}} - \tilde{V}_{n,l_{n,m}} \right|, \quad m = 1, \ldots, 2^N.
\]

This statement evidently follows from Lemmas 2.1 and 2.2 and from the relations (2.9)–(2.12), (2.22) and (2.23).

### 3 Proof of Theorem 2.1

In the proof of Theorem 2.1 we shall use the following auxiliary Lemmas 3.1–3.4 (Zaitsev 1995, 1996, 1998a).

**Lemma 3.1.** Suppose that \( \mathcal{L}(\xi) \in \mathcal{A}_d(\tau) \), \( y \in \mathbb{R}^m \), \( \alpha \in \mathbb{R}^1 \). Let \( M : \mathbb{R}^d \to \mathbb{R}^m \) be a linear operator and \( \tilde{\xi} \in \mathbb{R}^k \) be the vector consisting of a subset of coordinates of the vector \( \xi \). Then

\[
\mathcal{L}(M \xi + y) \in \mathcal{A}_m(\|M\| \tau), \quad \text{where} \quad \|M\| = \sup_{\|x\| \leq 1} \|Mx\|,
\]

\[
\mathcal{L}(\alpha \xi) \in \mathcal{A}_d(|\alpha| \tau), \quad \mathcal{L}(\tilde{\xi}) \in \mathcal{A}_k(\tau).
\]

**Lemma 3.2.** Suppose that independent random vectors \( \xi^{(k)} \), \( k = 1, 2 \), satisfy the condition \( \mathcal{L}(\xi^{(k)}) \in \mathcal{A}_{d_k}(\tau) \). Let \( \xi = \left( \xi^{(1)}, \xi^{(2)} \right) \in \mathbb{R}^{d_1+d_2} \) be the vector with the first \( d_1 \) coordinates coinciding with those of \( \xi^{(1)} \) and with the last \( d_2 \) coordinates coinciding with those of \( \xi^{(2)} \). Then \( \mathcal{L}(\xi) \in \mathcal{A}_{d_1+d_2}(\tau) \).

**Lemma 3.3.** (Bernstein-type inequality) Suppose that \( \mathcal{L}(\xi) \in \mathcal{A}_1(\tau) \), \( \mathbf{E} \xi = 0 \) and \( \mathbf{E} \xi^2 = \sigma^2 \). Then

\[
P\left\{ |\xi| \geq x \right\} \leq 2 \max \left\{ \exp \left( -x^2/4\sigma^2 \right), \exp \left( -x/4\tau \right) \right\}, \quad x \geq 0.
\]

**Lemma 3.4.** Let the distribution of a random vector \( \xi \in \mathbb{R}^d \) with \( \mathbf{E} \xi = 0 \) satisfy the condition \( \mathcal{L}(\xi) \in \overline{\mathcal{A}_d}(\tau, 4) \), \( \tau \geq 0 \). Assume that the variance \( \sigma^2 = \mathbf{E} \xi^2 > 0 \) of the last coordinate \( \xi_d \) of the vector \( \xi \) is the minimal eigenvalue of \( \text{cov} \xi \). Then there exist absolute positive constants \( c_8, \ldots, c_{12} \) such that the following assertions hold:
a) Let $d \geq 2$. Assume that $\xi_d$ is not correlated with previous coordinates $\xi_1, \ldots, \xi_{d-1}$ of the vector $\xi$. Define $B = \text{cov} \{P_{d-1} \xi\}$ and denote by $F(z \mid x)$, $z \in \mathbb{R}^1$, the r.c.d.f. of $\xi_d$ for a given value of $P_{d-1} \xi = x \in \mathbb{R}^{d-1}$. Let $L(P_{d-1} \xi) \in \mathcal{A}_d(\tau, 4)$. Then there exists $y \in \mathbb{R}^1$ such that
\[ |y| \leq c_8 \tau \left\| B^{-1/2} x \right\|^2 \leq c_8 \tau \left\| x \right\|^2 / \sigma^2, \] \hfill (3.1)
and
\[ \Phi_\sigma \left( z - \gamma(z) \right) < F(z + y \mid x) < \Phi_\sigma \left( z + \gamma(z) \right), \] \hfill (3.2)
for $\frac{\tau d^{3/2}}{\sigma} \leq c_9$, \[ \left| B^{-1/2} x \right| \leq \frac{c_{10} \sigma}{d^{3/2} \tau}, \] \[ |z| \leq \frac{c_{11} \sigma^2}{d \tau}, \]
where
\[ \gamma(z) = c_{12} \tau \left( d^{3/2} + d \delta \left( 1 + \frac{|z|}{\sigma} \right) + \frac{z^2}{\sigma^2} \right), \] \hfill (3.3)
\[ \delta = \left\| B^{-1/2} x \right\|. \]
b) The assertion a) remains valid for $d = 1$ with $F(z \mid x) = \mathbb{P} \{ \xi_1 < z \}$ and $y = \delta = 0$ without any restrictions on $B$, $P_{d-1} \xi$ and $x$.

**Remark 3.1.** In Zaitsev (1995, 1996) the formulation of Lemma 3.4 is in some sense weaker, see Zaitsev (1995, 1996, Lemmas 6.1 and 8.1). In particular, instead of the conditions
\[ L(\xi) \in \mathcal{A}_d(\tau, 4) \quad \text{and} \quad L(P_{d-1} \xi) \in \mathcal{A}_{d-1}(\tau, 4) \] \hfill (4.4)
the stronger conditions
\[ L(\xi) \in \mathcal{A}_d^*(\tau, 4, 4) \quad \text{and} \quad L(P_{d-1} \xi) \in \mathcal{A}_{d-1}^*(\tau, 4, 4) \] \hfill (4.5)
are used. However, in the proof of (3.1) and (3.2) only the conditions (3.4) are applied. The conditions (3.5) are necessary for the investigation of quantiles of conditional distributions corresponding to random vectors having coinciding moments up to third order which has been done in Zaitsev (1995, 1996) simultaneously with the proof of (3.1) and (3.2).

**Lemma 3.5.** Let $S_k = X_1 + \cdots + X_k$, $k = 1, \ldots, n$, be sums of independent random vectors $X_j \in \mathbb{R}^d$ and let $q(\cdot)$ be a semi-norm in $\mathbb{R}^d$. Then
\[ \mathbb{P} \left\{ \max_{1 \leq k \leq n} q(S_k) > 3t \right\} \leq 3 \max_{1 \leq k \leq n} \mathbb{P} \left\{ q(S_k) > t \right\}, \quad t \geq 0. \] \hfill (6.6)

Lemma 3.5 is a version of the Ottaviani inequality, see Dudley (1989, p. 251) or Hoffmann-Jørgensen (1994, p. 472). In the form (6.6) this inequality can be found in Etemadi (1985) with 4 instead of 3 (twice). The proof of Lemma 3.5 repeats those from the references above and is therefore omitted.

**Lemma 3.6.** Let the conditions of Theorem 2.1 be satisfied and assume that the vectors $X_k$, $k = 1, \ldots, 2^N$, are constructed by the dyadic procedure described in (2.22)–(2.36). Then there exist absolute positive constants $c_{13}, \ldots, c_{17}$ such that

a) If $\tau d^{3/2}/2^{N/2} \leq c_9$, then
\[ \left| U_{N,0} - V_{N,0} \right| \leq c_{13} d^{3/2} \tau \left( 1 + 2^{-N} \left| U_{N,0} \right|^2 \right) \] \hfill (7.7)
provided that $\left| U_{N,0} \right| \leq \frac{c_{14} 2^N}{d^{3/2} \tau}$;
b) If $1 \leq n \leq N$, $0 \leq k < 2^{N-n}$, $\tau d^{3/2}/2^{n/2} \leq c_{15}$, then

$$|\tilde{U}_{n,k} - \tilde{V}_{n,k}| \leq c_{16} d^{3/2} (1 + 2^{-n}) |U_{n,k}|^2$$  \hspace{1cm} (3.8)

provided that $|U_{n,k}| \leq \frac{c_{17} 2^n}{d^{3/2}}$.

In the proof of Lemma 3.6 we need the following auxiliary Lemma 3.7 which is useful for the application of Lemma 3.4 to the conditional distributions involved in the dyadic scheme.

**Lemma 3.7.** Let $F(\cdot)$ denote a continuous distribution function and $G(\cdot)$ an arbitrary distribution function satisfying for $z \in B \in \mathcal{B}_1$ the inequality

$$G(z - f(z)) < F(z + w) < G(z + f(z))$$

with some $f : B \to \mathbb{R}^1$ and $w \in \mathbb{R}^1$. Let $\eta \in \mathbb{R}^1$, $0 < G(\eta) < 1$ and $\xi = F^{-1}(G(\eta))$, where $F^{-1}(x) = \sup \left\{ u : F(u) \leq x \right\}$, $0 < x < 1$. Then

$$|\xi - \eta| < f(\xi - w) + |w|, \quad \text{if} \quad \xi - w \in B.$$  

**Proof**  Put $\zeta = \xi - w$. The continuity of $F$ implies that $F(F^{-1}(x)) \equiv x$, for $0 < x < 1$. Therefore,

$$\zeta \in B \Rightarrow G(\zeta - f(\zeta)) < F(\xi) = G(\eta) \Rightarrow \zeta - f(\zeta) < \eta \Rightarrow \xi - \eta < f(\zeta) + w$$

and

$$\zeta \in B \Rightarrow G(\eta) = F(\xi) < G(\zeta + f(\zeta)) \Rightarrow \eta < \zeta + f(\zeta) \Rightarrow \eta - \xi < f(\zeta) - w.$$  

This completes the proof of the lemma. □

**Proof of Lemma 3.6**  At first we note that the conditions of Theorem 2.1 imply that

$$\text{cov} \ U_{n,k} = 2^n I_{2d}. \quad \text{for} \quad 1 \leq n \leq N, \quad 0 \leq k < 2^{N-n},$$

and, hence (see (2.28)),

$$\text{cov} \ U_{n,k}^j = 2^n I_j, \quad \text{for} \quad 1 \leq j \leq 2d.$$

(3.9)

Let us prove the assertion a). Introduce the vectors

$$U_{N,0}^j = \left( U_{N,0}^{(1)}, \ldots, U_{N,0}^{(j)} \right), \quad V_{N,0}^j = \left( V_{N,0}^{(1)}, \ldots, V_{N,0}^{(j)} \right)$$

consisting of the first $j$ coordinates of the vectors $U_{N,0}, V_{N,0}$ respectively. By (3.9), (2.32) and (2.34),

$$U_{N,0} = \overline{P}_d \ U_{N,0}$$

(3.11)

and

$$U_{N,0}^j = U_{N,0}^j, \quad \text{cov} \ U_{N,0}^j = 2^n I_j, \quad \text{for} \quad 1 \leq j \leq d.$$  

(3.12)
Moreover, according to Lemma 2.1, Remark 2.1 (3.12) and (2.19), the distributions \( \mathcal{L}(U_{N,0}^j) \), \( j = 1, \ldots, d \), satisfy in the \( j \)-dimensional case the conditions of Lemma 3.4 with \( \sigma^2 = 2^N \) and \( B = \text{cov} U_{N,0}^{-1} = 2^N I_{j-1} \) (the last equality for \( j \geq 2 \)).

Taking into account (2.29) and applying Lemmas 3.4 and 3.7, we see that one can choose \( c_2, c_3 \), (3.10), and \( \tau \), (3.12), (2.19), (2.26) to be so small and \( \tau \leq c_9 \), \( |U_{N,0}^{(1)}| \leq c_{11} \frac{2^N}{\tau} \). Furthermore,

\[
|U_{N,0}^{(j)} - V_{N,0}^{(j)}| \leq c_{12} \tau \left( j^{3/2} + j^{3/2} \frac{|U_{N,0}^{j-1}|}{2^{N/2}} \left( 1 + \frac{|U_{N,0}^{j-1} - y_j|}{2^{N/2}} \right) + \frac{|U_{N,0}^{j-1} - y_j|^2}{2^N} \right) + |y_j|, \tag{3.14}
\]

if \( \frac{j^{3/2}}{2^{N/2}} \leq c_9 \), \( \frac{|U_{N,0}^{j-1}|}{2^{N/2}} \leq \frac{c_{10} \cdot 2^{N/2}}{j^{3/2} \tau} \), \( |U_{N,0}^{(j)} - y_j| \leq \frac{c_{11} \cdot 2^N}{j \tau} \), \( 2 \leq j \leq d \), (3.15)

where

\[
|y_j| \leq c_{12} \tau j \frac{|U_{N,0}^{j-1}|^2}{2^N}, \quad 2 \leq j \leq d. \tag{3.16}
\]

Obviously,

\[
|U_{N,0}^{(1)}| \leq \max \left\{ \left| U_{N,0}^{j-1} \right|, \left| U_{N,0}^{(j)} \right| \right\} = \left| U_{N,0}^{j-1} \right| \leq \left| U_{N,0} \right|, \quad 2 \leq j \leq d, \tag{3.17}
\]

see (3.11) and (3.10). Using (3.13), (3.14), (3.16) and (3.17), we see that one can choose \( c_{13} \) to be so small and \( c_{14} \) to be so small that

\[
|U_{N,0}^{(j)} - V_{N,0}^{(j)}| \leq c_{13} d^{3/2} \tau \left( 1 + 2^{-N} \left| U_{N,0} \right|^2 \right), \tag{3.18}
\]

if \( \frac{d^{3/2}}{2^{N/2}} \leq c_9 \), \( |U_{N,0}| \leq c_{14} \frac{2^N}{d^{3/2} \tau} \), \( 1 \leq j \leq d \). The inequality (3.7) immediately follows from (3.18), (2.23) and (2.31).

Now we shall prove item b). According to Lemma 2.1, Remark 2.1 (2.18), (2.31) and (3.9), the distributions \( \mathcal{L}(U_{n,k}^j) \), \( j = d + 1, \ldots, 2d \), satisfy in the \( j \)-dimensional case the conditions of Lemma 3.4 with \( \sigma^2 = 2^n \), \( B = \text{cov} U_{n,k}^{-1} = 2^n I_{j-1} \).

Using (2.29) and applying Lemmas 3.4 and 3.7, we see that one can choose \( c_2, c_3 \), (3.10), (3.12), (2.19), (2.26) to be so small and \( \tau \leq c_9 \), \( |U_{n,k}^{(1)}| \leq c_{11} \frac{2^n}{\tau} \). Furthermore,

\[
|U_{n,k}^{(j)} - V_{n,k}^{(j)}| \leq c_{12} \tau \left( j^{3/2} + j^{3/2} \frac{|U_{n,k}^{j-1}|}{2^{n/2}} \left( 1 + \frac{|U_{n,k}^{j-1} - y_j|}{2^{n/2}} \right) + \frac{|U_{n,k}^{j-1} - y_j|^2}{2^n} \right) + |y_j|, \tag{3.19}
\]

if \( \frac{j^{3/2}}{2^{n/2}} \leq c_9 \), \( \frac{|U_{n,k}^{j-1}|}{2^{n/2}} \leq \frac{c_{10} \cdot 2^{n/2}}{j^{3/2} \tau} \), \( |U_{n,k}^{(j)} - y_j| \leq \frac{c_{11} \cdot 2^n}{j \tau} \), \( 1 \leq j \leq 2d \), (3.15)
If

\[ \frac{\tau_{j}^{3/2}}{2^{n/2}} \leq c_9, \quad \frac{|U_{n,k}^{j-1}|}{2^{n/2}} \leq c_{10} \cdot \frac{2^{n/2}}{j^{3/2}}, \quad \frac{|U_{n,k}^{(j)} - y_j|}{2^n} \leq \frac{c_{11} \cdot 2^n}{j \tau}, \]

(3.20)

where

\[ |y_j| \leq c_8 \tau_j \frac{|U_{n,k}^{j-1}|^2}{2^n}, \quad d + 1 \leq j \leq 2d. \]

(3.21)

Obviously,

\[ \max \left\{ \left| U_{n,k}^{j-1} \right|, \left| U_{n,k}^{(j)} \right| \right\} = \left| U_{n,k}^{j} \right| \leq \left| U_{n,k} \right|, \]

(3.22)

see (2.34). Using (3.19), (3.21) and (3.22), we see that one can choose \( c_{15} \) and \( c_{17} \) to be so small and \( c_{16} \) to be so large that

\[ \left| U_{n,k}^{(j)} - V_{n,k}^{(j)} \right| \leq c_{16} d^{3/2} \tau \left( 1 + 2^{-n} \left| U_{n,k} \right|^2 \right) \]

(3.23)

if \( \frac{\tau_{d}^{3/2}}{2^{n/2}} \leq c_{15}, \left| U_{n,k} \right| \leq \frac{c_{17} \cdot 2^n}{d^{3/2} \tau}, \quad d + 1 \leq j \leq 2d. \) The inequality (3.8) immediately follows from (3.23), (2.24), (2.25) and (2.34). □

**Proof of Theorem 2.1** Let \( X_k, \quad k = 1, \ldots, 2^N \), denote the vectors constructed by the dyadic procedure described in (2.22)–(2.36). Denote

\[ \Delta = \Delta(2^N) = \max_{1 \leq k \leq 2^N} \left| S_k - T_k \right|, \]

(3.24)

\[ c_5 = \min \left\{ c_9, c_{15} \right\}, \quad c_{18} = \min \left\{ c_{14}, c_{17}, 1 \right\}, \quad y := \frac{c_{18}}{d^{3/2} \tau} \leq \frac{1}{\tau}, \]

(3.25)

fix some \( x > 0 \) and choose the integer \( M \) such that

\[ x < 4y \cdot 2^M \leq 2x. \]

(3.26)

We shall estimate \( P \left\{ \Delta \geq x \right\} \). Consider separately two possible cases: \( M \geq N \) and \( M < N \). Let, at first, \( M \geq N \). Denote

\[ \Delta_1 = \max_{1 \leq k \leq 2^N} \left| S_k \right|, \quad \Delta_2 = \max_{1 \leq k \leq 2^N} \left| T_k \right|. \]

(3.27)

It is easy to see that \( \Delta \leq \Delta_1 + \Delta_2 \) and, hence,

\[ P \left\{ \Delta \geq x \right\} \leq P \left\{ \Delta_1 \geq x/2 \right\} + P \left\{ \Delta_2 \geq x/2 \right\}. \]

(3.28)

Taking into account the completeness of classes \( A_d(\tau) \) with respect to convolution, applying Lemmas 3.5, 3.1 and 3.3 and using (3.25) and (3.26), we obtain that \( 2^N \leq 2^M \leq x/2y \) and

\[ P \left\{ \Delta_1 \geq x/2 \right\} \leq 3 \max_{1 \leq k \leq 2^N} P \left\{ \left| S_k \right| \geq x/6 \right\} \]

\[ \leq 6d \exp \left( -\min \left\{ \frac{x^2}{144 \cdot 2^N}, \frac{x}{24 \tau} \right\} \right) \]

\[ \leq 6d \exp \left( -\frac{c_{19} x}{d^{3/2} \tau} \right). \]

(3.29)
Since all $d$-dimensional Gaussian distributions belong to all classes $A_d(\tau)$, $\tau \geq 0$, we automatically obtain that

$$\mathbb{P}\left\{ \Delta_2 \geq x/2 \right\} \leq 6d \exp\left( -\frac{c_{19}x}{d^{3/2} \tau} \right).$$  \hspace{1cm} (3.30)

From (3.28)–(3.30) it follows in the case $M \geq N$ that

$$\mathbb{P}\left\{ \Delta \geq x \right\} \leq 12d \exp\left( -\frac{c_{19}x}{d^{3/2} \tau} \right).$$  \hspace{1cm} (3.31)

Let now $M < N$. Denote

$$L = \max\left\{ 0, M \right\}$$  \hspace{1cm} (3.32)

and

$$\Delta_3 = \max_{0 \leq k < 2^{N-L}} \max_{1 \leq l \leq 2L} \left| S_{k, 2^L+l} - S_{k, 2^L} \right|,$$  \hspace{1cm} (3.33)

$$\Delta_4 = \max_{0 \leq k < 2^{N-L}} \max_{1 \leq l \leq 2L} \left| T_{k, 2^L+l} - T_{k, 2^L} \right|,$$  \hspace{1cm} (3.34)

$$\Delta_5 = \max_{1 \leq k \leq 2^{N-L}} \left| S_{k, 2^L} - T_{k, 2^L} \right|.$$  \hspace{1cm} (3.35)

Introduce the event

$$A = \left\{ \omega : \left| U_{L,k} \right| < y \cdot 2^L, \ 0 \leq k < 2^{N-L} \right\}$$  \hspace{1cm} (3.36)

(we assume that all considered random vectors are measurable mappings of $\omega \in \Omega$). For the complementary event we use the notation $\overline{A} = \Omega \setminus A$.

We consider separately two possible cases: $L = M$ and $L = 0$. Let $L = M$. It is evident that in this case

$$\Delta \leq \Delta_3 + \Delta_4 + \Delta_5.$$  \hspace{1cm} (3.37)

Moreover, by virtue of (3.37), (3.26), (3.33) and (3.36), we have

$$\overline{A} \subset \left\{ \omega : \Delta_3 \geq x/4 \right\}.$$  \hspace{1cm} (3.38)

From (3.37) and (3.38) it follows that

$$\mathbb{P}\left\{ \Delta \geq x \right\} \leq \mathbb{P}\left\{ \Delta_3 \geq x/4 \right\} + \mathbb{P}\left\{ \Delta_4 \geq x/4 \right\} + \mathbb{P}\left\{ \Delta_5 \geq x/2, A \right\}.$$  \hspace{1cm} (3.39)

Using Lemmas 3.3, 3.1 and 3.3, the completeness of classes $A_d(\tau)$ with respect to convolution and the relations (3.25) and (3.26), we obtain, for $0 \leq k < 2^{N-L}$, that $2^L = 2^M \leq x/2y$ and

$$\mathbb{P}\left\{ \max_{1 \leq l \leq 2L} \left| S_{k, 2^L+l} - S_{k, 2^L} \right| \geq x/4 \right\} \leq 3 \max_{1 \leq l \leq 2L} \mathbb{P}\left\{ \left| S_{k, 2^L+l} - S_{k, 2^L} \right| \geq x/12 \right\}$$

$$\leq 6d \exp\left( -\min\{\frac{x^2}{576 \cdot 2^L}, \frac{x}{48 \tau} \} \right)$$

$$\leq 6d \exp\left( -\frac{c_{20}x}{d^{3/2} \tau} \right).$$  \hspace{1cm} (3.40)

Since all $d$-dimensional Gaussian distributions belong to classes $A_d(\tau)$ for all $\tau \geq 0$, we immediately obtain that

$$\mathbb{P}\left\{ \max_{1 \leq l \leq 2L} \left| T_{k, 2^L+l} - T_{k, 2^L} \right| \geq x/4 \right\} \leq 6d \exp\left( -\frac{c_{20}x}{d^{3/2} \tau} \right).$$  \hspace{1cm} (3.41)
From (3.33), (3.34), (3.40) and (3.41) it follows that
\[
P \left\{ \Delta_3 \geq x/4 \right\} + P \left\{ \Delta_4 \geq x/4 \right\} \leq 2^N \cdot 12d \exp \left( - \frac{c_{20}x}{d^{3/2} \tau} \right). \tag{3.42}
\]

Assume that \( L = 0 \). Then, according to (3.24) and (3.35), \( \Delta = \Delta_5 \) and, hence, we have the rough bound
\[
P \left\{ \Delta \geq x \right\} \leq P \left\{ \mathcal{T} \right\} + P \left\{ \Delta_5 \geq x/2, A \right\}. \tag{3.43}
\]
In this case \( U_{L,k} = X_{k+1}, 2^L = 1 \geq 2^M, \ y > x/4 \) (see (3.25), (3.26) and (3.32)). Therefore, by (3.36) and by Lemmas 3.1 and 3.3,
\[
P \left\{ \mathcal{T} \right\} \leq \sum_{k=0}^{2^{N-1}} P \left\{ \left| U_{L,k} \right| \geq y \cdot 2^L \right\} = \sum_{k=1}^{2^N} P \left\{ \left| X_k \right| \geq y \right\}
\leq 2^{N+1} d \exp \left( - \min \left\{ \frac{y^2}{4}, \frac{y}{4\tau} \right\} \right)
\leq 2^{N+1} d \exp \left( - \min \left\{ \frac{xy}{16}, \frac{x}{16 \tau} \right\} \right)
\leq 2^{N+1} d \exp \left( - \frac{c_{21}x}{d^{3/2} \tau} \right). \tag{3.44}
\]

It remains to estimate \( P \left\{ \Delta_5 \geq x/2, A \right\} \) in both cases: \( L = M \) and \( L = 0 \) (see (3.39) and (3.42)–(3.44)). Let \( L \) defined by (3.32) be arbitrary. Fix an integer \( k \) satisfying \( 1 \leq k \leq 2^{N-L} \) and denote for simplicity
\[
j = j(k) := k \cdot 2^L. \tag{3.45}
\]
By Corollary 2.1 we have
\[
\left| S_{k,2^L} - T_{k,2^L} \right| = \left| S_j - T_j \right| \leq \left| U_{N,0} - V_{N,0} \right| + \frac{1}{2} \sum_{n=L+1}^{N} \left| \tilde{U}_{n,l_{n,j}} - \tilde{V}_{n,l_{n,j}} \right|, \tag{3.46}
\]
where \( l_{n,j} \) are integers, defined by \( l_{n,j} \cdot 2^n < j \leq (l_{n,j} + 1) \cdot 2^n \) (see (2.42)).

By virtue of (3.25) and (3.36), for \( \omega \in A \) we have
\[
\left| U_{L,l} \right| < y \cdot 2^L = \frac{c_{18} \cdot 2^L}{d^{3/2} \tau} \leq \frac{\min\{c_{14}, c_{17}\} \cdot 2^L}{d^{3/2} \tau}, \quad 0 \leq l < 2^{N-L}, \tag{3.47}
\]
and, by (3.25)–(3.37), \( U_{L,l} \) are sums over blocks consisting of \( 2^L \) summands. Moreover, \( U_{n,l} \) (resp. \( \tilde{U}_{n,l} \)), \( L + 1 \leq n \leq N, 0 \leq l < 2^{N-n} \), are sums (resp. differences) of two sums over blocks containing each \( 2^{n-1} \) summands. These sums and differences can be represented as linear combinations (with coefficients \( \pm 1 \)) of \( 2^{n-L} \) sums over blocks containing each \( 2^L \) summands and satisfying (3.47). Therefore, for \( \omega \in A, L + 1 \leq n \leq N, 0 \leq l < 2^{N-n} \) we have (see (2.32) and (2.34))
\[
\left| U_{n,l} \right| = \max \left\{ \left| U_{n,l} \right|, \left| \tilde{U}_{n,l} \right| \right\} \leq 2^{n-L} y \cdot 2^L = y \cdot 2^n \leq \frac{\min\{c_{14}, c_{17}\} \cdot 2^n}{d^{3/2} \tau}. \tag{3.48}
\]
Using (3.48), we see that if \( \omega \in A \), the conditions of Lemma 3.6 are satisfied for \( \tau, U_{N,0} \) and \( U_{n,l} \), if \( L + 1 \leq n \leq N, 0 \leq l < 2^{N-n} \). By (3.46), (3.48) and by Lemma 3.6 for \( \omega \in A \) we have

\[
|S_j - T_j| \leq c_{13} d^{3/2} \tau \left( 1 + 2^{-N} \left| U_{N,0} \right|^2 \right)
+ \sum_{n=L+1}^{N} c_{16} d^{3/2} \tau \left( 1 + 2^{-n} \max\left\{ \left| U_{n,l_{n,j}} \right|^2, \left| \tilde{U}_{n,l_{n,j}} \right|^2 \right\} \right)
\leq c d^{3/2} \tau \left( N + 1 + 2^{-N} \left| U_{N,0} \right|^2 + \sum_{n=L}^{N-1} 2^{-n} \left( \left| U^{(n)} \right|^2 + \left| U_{(n)} \right|^2 \right) \right),
\]

(3.49)

where

\[
U^{(n)} = U_{n,l_{n,j}}, \quad U_{(n)} = U_{n,l_{n,j}};
\]

and

\[
\tilde{l}_{n-1,j} = \begin{cases} 
2l_{n,j}, & \text{if } l_{n-1,j} = 2l_{n,j} + 1, \\
2l_{n,j} + 1, & \text{if } l_{n-1,j} = 2l_{n,j}, \quad L < n \leq N
\end{cases}
\]

(3.50)

(3.51)

(it is easy to see that \( l_{n-1,j} \) can be equal either to \( 2l_{n,j} \) or to \( 2l_{n,j} + 1 \), for given \( l_{n,j} \)). In other words, \( U^{(n)}, L \leq n \leq N, \) is the sum over the block of \( 2^n \) summands which contains \( X_j \). The sum \( U_{(n)} \) does not contain \( X_j \) and

\[
U^{(n+1)} = U^{(n)} + U_{(n)}, \quad L \leq n < N
\]

(3.52)

(see (3.37)). The equality (3.52) implies

\[
U^{(n)} = U^{(L)} + \sum_{s=0}^{n-L-1} U_{(L+s)}, \quad L \leq n \leq N.
\]

(3.53)

It is important that all summands in the right-hand side of (3.53) are the sums of disjoint blocks of independent summands. Therefore, they are independent.

Put \( \beta = 1/\sqrt{2} \). Then, using (3.53) and the Hölder inequality, one can easily derive that, for \( L \leq n \leq N, \)

\[
\left| U^{(n)} \right|^2 \leq c_{22} \left( \beta^{-(n-L)} \left| U^{(L)} \right|^2 + \sum_{s=0}^{n-L-1} \beta^{-(n-L-1)+s} \left| U_{(L+s)} \right|^2 \right),
\]

(3.54)

with \( c_{22} = \sum_{j=0}^{\infty} \beta^j = \frac{\sqrt{2}}{\sqrt{2} - 1} \). It is easy to see that

\[
\sum_{n=L}^{N} 2^{-n} \beta^{-(n-L)} \left| U^{(L)} \right|^2 \leq c_{22} \cdot 2^{-L} \left| U^{(L)} \right|^2.
\]

(3.55)

Moreover,

\[
\sum_{n=L+1}^{N} \sum_{s=0}^{n-L-1} 2^{-n} \beta^{-(n-L-1)+s} \left| U_{(L+s)} \right|^2
\]
\[
\begin{align*}
&= \sum_{s=0}^{N-L-1} \sum_{n=L+1+s}^{N} 2^{-n} \beta^{-(n-L-1)+s} \left| U_{(L+s)} \right|^2 \\
&\leq c_{22} \sum_{s=0}^{N-L-1} 2^{-(L+1+s)} \left| U_{(L+s)} \right|^2 . \quad (3.56)
\end{align*}
\]

It is clear that the inequalities (3.54)–(3.56) imply
\[
2^{-N} \left| U_{N,0} \right|^2 + \sum_{n=L}^{N-1} 2^{-n} \left( \left| U^{(n)} \right|^2 + \left| U_{(n)} \right|^2 \right) \\
\leq c_{22} \left( \frac{\left| U^{(L)} \right|^2}{2^L} + \sum_{s=0}^{N-L-1} \frac{\left| U_{(L+s)} \right|^2}{2^{L+1+s}} \right) + \sum_{n=L}^{N-1} \frac{\left| U_{(n)} \right|^2}{2^n} \\
\leq c \left( \frac{\left| U^{(L)} \right|^2}{2^L} + \sum_{n=L}^{N-1} \frac{\left| U_{(n)} \right|^2}{2^n} \right). \quad (3.57)
\]

From (3.49) and (3.57) it follows that for \( \omega \in A \) we have
\[
\left| S_j - T_j \right| \leq c_{23} d^{3/2} \tau \left( N + 1 + \frac{\left| U^{(L)} \right|^2}{2^L} + \sum_{n=L}^{N-1} \frac{\left| U_{(n)} \right|^2}{2^n} \right). \quad (3.58)
\]

Denote (for \( 0 \leq n \leq N, \ 0 \leq l < 2^{N-n} \))
\[
W_{n,l} = \begin{cases} 
2^{-n} \left| U_{n,l} \right|^2, & \text{if } \left| U_{n,l} \right| \leq y \cdot 2^n, \\
0, & \text{otherwise.}
\end{cases} \quad (3.59)
\]

Let us show that
\[
\mathbf{E} \exp \left( t W_{n,l} \right) \leq 2d + 1, \quad \text{for } 0 \leq t \leq \frac{1}{8}. \quad (3.60)
\]

Indeed, integrating by parts, we obtain
\[
\mathbf{E} \exp(t W_{n,l}) = 1 + \int_{0}^{y^2 \cdot 2^n} t \exp(tu) \, \mathbf{P} \left\{ W_{n,l} \geq u \right\} \, du \\
\leq 1 + \frac{1}{8} \int_{0}^{y^2 \cdot 2^n} \exp \left( u/8 \right) \, \mathbf{P} \left\{ \left| U_{n,l} \right| \geq 2^n \sqrt{u} \right\} \, du. \quad (3.61)
\]

Taking into account (3.37), (3.25) and using Lemmas 3.1 and 3.3, we obtain that
\[
\mathbf{P} \left\{ \left| U_{n,l} \right| \geq 2^n \sqrt{u} \right\} \leq 2d \exp \left( - \min \left\{ \frac{2^n u}{4 \cdot 2^n}, \frac{2^n \sqrt{u}}{4 \tau} \right\} \right)
\]
\[ \leq 2d \exp \left( -\min \left\{ \frac{u}{4}, \frac{u}{4y\tau} \right\} \right) \]

\[ = 2d \exp \left( -\frac{u}{4} \right), \quad (3.62) \]

if \( 0 \leq u \leq y^2 \cdot 2^n \). The relation (3.60) immediately follows from (3.61) and (3.62).

The relations (3.47), (3.48) and (3.59) imply that, for \( L \leq n \leq N, \ 0 \leq l < 2^{N-n}, \ \omega \in A, \)

\[ 2^{-n} \left| U_{n,l} \right|^2 = W_{n,l}. \quad (3.63) \]

Thus, according to (3.50), we can rewrite (3.58) in the form

\[ \left| S_j - T_j \right| \leq c_{23} d^{3/2} \left( N + 1 + W^{(L)} + \sum_{n=L}^{N-1} W_{(n)} \right), \quad \omega \in A, \quad (3.64) \]

where

\[ W^{(L)} = W_{L,L,j}, \quad W_{(n)} = W_{n,\tilde{l}_{n,j}}. \quad (3.65) \]

Putting now \( t^* = \left( 8 c_{23} d^{3/2} \tau \right)^{-1} \) and \( t = t^* \cdot c_{23} d^{3/2} \tau = 1/8 \), taking into account that the random variables \( W^{(L)}, W_{(L)}, \ldots, W_{(N-1)} \) are independent and applying (3.60), (3.64) and (3.65), we obtain

\[
\mathbb{P}\left\{ \left\{ \omega : \left| S_j - T_j \right| \geq x/2 \right\} \cap A \right\} \\
\leq \mathbb{P}\left\{ c_{23} d^{3/2} \left( N + 1 + W^{(L)} + \sum_{n=L}^{N-1} W_{(n)} \right) \geq x/2 \right\} \\
\leq \mathbb{P}\left\{ t \left( W^{(L)} + \sum_{n=L}^{N-1} W_{(n)} \right) \geq t^* x/2 - t(N + 1) \right\} \\
\leq \mathbb{E} \exp\left( t \left( W^{(L)} + \sum_{n=L}^{N-1} W_{(n)} \right) \right) \exp\left( t^* x/2 - t(N + 1) \right) \\
= \mathbb{E} \exp\left( t W^{(L)} \right) \prod_{n=L}^{N-1} \mathbb{E} \exp\left( t W_{(n)} \right) \exp\left( t^* x/2 - t(N + 1) \right) \\
\leq (3d)^{N+1} \exp\left( \frac{N + 1}{8} - \frac{x}{16 c_{23} d^{3/2} \tau} \right). \quad (3.66) \]

From (3.35), (3.45) and (3.66) it follows that

\[
\mathbb{P}\left\{ \Delta_5 \geq x/2, \ A \right\} \leq 2^N \cdot (3d)^{N+1} \exp\left( \frac{N + 1}{8} - \frac{x}{16 c_{23} d^{3/2} \tau} \right). \quad (3.67) \]

Using (3.31), (3.39), (3.42)–(3.44) and (3.67), we obtain that

\[
\mathbb{P}\left\{ \Delta \geq x \right\} \leq (19d)^{N+1} \exp\left( -\frac{x}{c_{24} d^{3/2} \tau} \right), \quad x \geq 0, \quad (3.68) \]
where we can take \( c_{24} = \max \{ 16 c_{23}, c_{19}^{-1}, c_{20}^{-1}, c_{21}^{-1}, 2 \} \). Let the quantities \( \varepsilon, x_0 > 0 \) be defined by the relations

\[
\varepsilon = \frac{1}{2c_{24}d^{3/2}\tau} \leq \frac{1}{4\tau}, \quad e^{\varepsilon x_0} = (19d)^{N+1}. \tag{3.69}
\]

Integrating by parts and using (3.68) and (3.69), we obtain

\[
\mathbb{E} e^{\varepsilon \Delta} = \int_0^\infty \varepsilon e^{\varepsilon x} P \left\{ \Delta \geq x \right\} dx + 1,
\]

\[
\int_0^{x_0} \varepsilon e^{\varepsilon x} P \left\{ \Delta \geq x \right\} dx \leq \int_0^{x_0} \varepsilon e^{\varepsilon x} dx = e^{\varepsilon x_0} - 1 = (19d)^{N+1} - 1,
\]

\[
\int_{x_0}^\infty \varepsilon e^{\varepsilon x} P \left\{ \Delta \geq x \right\} dx \leq \int_{x_0}^\infty \varepsilon e^{-\varepsilon (x-x_0)} dx = 1,
\]

and, hence,

\[
\mathbb{E} e^{\varepsilon \Delta} \leq (19d)^{N+1} + 1 \leq (20d)^{N+1}.
\]

Together with (3.24) and (3.69), this completes the proof of Theorem 2.1. \( \square \)

### 4 Proofs of Theorems 1.1–1.4

We start the proofs of Theorems 1.1–1.3 with the following common part.

**Beginning of the proofs of Theorems 1.1, 1.2 and 1.3** At first we shall verify that under the conditions of Theorems 1.2 or 1.3 we have \( \mathcal{L}(\xi_k) \in \mathcal{A}_d(\tau) \). For Theorem 1.3 this relation is an immediate consequence of Lemma 3.1, of the completeness of classes \( \mathcal{A}_d(\tau) \) with respect to convolution and of the conditions (1.8) and (1.10)–(1.12). In the case of Theorem 1.2 we denote \( K = \mathcal{L}(\eta) \). One can easily verify that \( B = \text{cov} K = \gamma^2 I_d \), where \( \gamma^2 \) is defined by (1.7) and, hence,

\[
1 \leq \gamma^2 \leq 3. \tag{4.1}
\]

Moreover,

\[
\varphi(K, z) = \log \mathbb{E} e^{\langle z, \eta \rangle} = \log \left( \frac{4 + \tau^2 (d + \langle z, \tau \rangle^2)}{4 + \tau^2 d} \right) \exp \left( \frac{\langle z, \tau \rangle^2}{2} \right), \quad z \in \mathbb{C}^d. \tag{4.2}
\]

Using (4.1) and (4.2), we obtain

\[
\left| d_u d_v \varphi(K, z) \right| = \left| d_u d_v \log \left( 4 + \tau^2 (d + \langle z, \tau \rangle^2) \right) \right| \leq c \tau^3 \| u \| \| v \|^2 \leq \| u \| \| v \| \tau \left\langle B v, v \right\rangle, \tag{4.3}
\]

for \( \| z \| \tau \leq 1 \), provided that \( c_1 \) (involved in Assertion A) is sufficiently small. This means that \( K = \mathcal{L}(\eta) \in \mathcal{A}_d(\tau) \). The relation \( \mathcal{L}(\xi_k) = \mathcal{L}(\eta/\gamma) \in \mathcal{A}_d(\tau), \ k = 1, \ldots, n \), follows from (4.1) and from Lemma 3.1.
The text below is related to Theorems 1.1, 1.2 and 1.3 simultaneously. Without loss of generality we assume that the amount of summands is equal to \( 2^N \) with some positive integer \( N \). It suffices to show that the dyadic scheme related to the vectors \( \xi_1, \ldots, \xi_{2^n} \) satisfies the conditions of Theorem 2.1 with \( \tau^* = \sqrt{2} \tau \) instead of \( \tau \). According to Lemma 2.1 we can verify the conditions (2.18) and (2.19) for the vectors \( u_{n,k} \) instead of \( U_{n,k} \) and \( U_{n,0} \). To this end we shall show that

\[
\mathcal{L}(U_{n,k}) \in \mathcal{A}_j(\sqrt{2} \tau, 4) \quad \text{for} \quad 0 \leq k < 2^{N-n}, \quad 1 \leq n \leq N, \quad 1 \leq j \leq 2d. \tag{4.4}
\]

Recall that \( U_{n,k} = A \tilde{U}_{n,k} \) where \( A \) is the linear operator defined by (2.16) and satisfying (2.40). Furthermore, \( \tilde{U}_{n,k} = (U_{n,1,2k}, U_{n,1,2k+1}) \in \mathbb{R}^{2d} \), where the \( d \)-dimensional vectors \( U_{n,1,2k} \) and \( U_{n,1,2k+1} \) are independent. The relation \( \mathcal{L}(U_{n,k}) \in \mathcal{A}_{2d}(\sqrt{2} \tau) \) can be therefore easily derived from the conditions of Theorems 1.1 and 1.2 with the help of Lemmas 2.1, 3.1 and 3.2 (see (2.40)) if we take into account the completeness of classes \( \mathcal{A}_d(\tau) \) with respect to convolution and their monotonicity with respect to \( \tau \). It is easy to see that \( U_{n,k} = P_j U_{n,k} \), where the projector \( P_j : \mathbb{R}^{2d} \to \mathbb{R}^j \) can be considered as a linear operator with \( \| P_j \| = 1 \) (see (2.34)). Applying Lemma 3.1 again, we obtain the relations \( \mathcal{L}(U_{n,k}) \in \mathcal{A}_j(\sqrt{2} \tau), 1 \leq j \leq 2d \).

It remains to verify that, for \( h \in \mathbb{R}^j \), \( \| h \| \sqrt{2} \tau < 1 \), the following inequality hold:

\[
\int_T \left| \hat{F}_h(t) \right| dt \leq \frac{(2\pi)^{j/2} \sqrt{2} \tau j^{3/2}}{\sigma (\det D)^{1/2}}, \tag{4.5}
\]

\[
T = \left\{ t \in \mathbb{R}^j : 4 \| t \| \sqrt{2} \tau j \geq 1 \right\}, \tag{4.6}
\]

where \( F = \mathcal{L}(U_{n,k}) \), and \( \sigma^2 \) is the minimal eigenvalue of \( D = \text{cov} U_{n,k} \). Note that, according to (3.9), we have

\[
D = 2^n I_j, \quad \sigma^2 = 2^n, \quad \det D = 2^{nj}. \tag{4.7}
\]

Introduce \( 2^{n-1} \) random vectors

\[
X_r = (X_r, X_{2^{n-1}+r}) \in \mathbb{R}^{2d}, \quad r = 2^{n-1} \cdot 2k + 1, \ldots, 2^{n-1}(2k + 1). \tag{4.8}
\]

Obviously, these vectors are independent. According to (2.36), (4.37) and (4.8),

\[
\tilde{U}_{n,k} = (U_{n,1,2k}, U_{n,1,2k+1}) = \sum_{r=2^{n-1}-2k+1}^{2^n(2k+1)} X_r. \tag{4.9}
\]

Denote now \( R^{(s)}_h = \mathcal{L}(X_s)(h) \), for \( s = 1, \ldots, 2^N \), \( h \in \mathbb{R}^d \), and \( M^{(r)}_h := \mathcal{L}(X_r)(h) \), \( Q^{(r)}_h := \mathcal{L}(AX_r)(h) \), for \( r = 2^{n-1} \cdot 2k + 1, \ldots, 2^{n-1}(2k + 1) \), \( h \in \mathbb{R}^{2d} \). As usually, we consider only such \( h \) for which these distributions exist. Using (2.8), we see that, for all \( t \in \mathbb{R}^{2d} \),

\[
\tilde{Q}^{(r)}_h(t) = \frac{E \exp \left( \left\langle h + it, AX_r \right\rangle \right)}{E \exp \left( \left\langle h, AX_r \right\rangle \right)} = \frac{E \exp \left( \left\langle A^*h + iA^*t, X_r \right\rangle \right)}{E \exp \left( \left\langle A^*h, X_r \right\rangle \right)} = \tilde{M}^{(r)}_{A^*h}(A^*t). \tag{4.10}
\]
By (2.3) and (4.9), we have (for \( j = 2d \))
\[
\left| \hat{F}_h(t) \right| = \prod_{r=2^{n-1}(2k+1)}^{2^{n-1}(2k+1)} \left| \hat{Q}^{(r)}(t) \right|.
\]
(4.11)

Split \( t = (t_1, \ldots, t_{2d}) \in \mathbb{R}^{2d} \) as \( t = (t^{(1)}, t^{(2)}) \), where we denote \( t^{(1)} = (t_1, \ldots, t_d) \) and \( t^{(2)} = (t_{d+1}, \ldots, t_{2d}) \in \mathbb{R}^d \). Using formulae (2.8) and (4.8) and introducing a similar notation for \( h \in \mathbb{R}^{2d} \), it is easy to check that
\[
\hat{M}_h^{(r)}(t) = \hat{R}^{(r)}_{h^{(1)}}(t^{(1)}) \hat{R}^{(2^{n-1}+r)}_{h^{(2)}}(t^{(2)}).
\]
(4.12)

Note that
\[
\| t \|^2 = \| t^{(1)} \|^2 + \| t^{(2)} \|^2.
\]
(4.13)

**End of the proof of Theorem 1.1.** Let now the distributions \( \mathcal{L}(\xi) \) satisfy the conditions of Theorem 1.1. In this case, according to (2.3), we have \( \hat{R}_h^{(s)} = H_s(h) \hat{G}(h) \). It is well-known that the conjugate distributions \( \hat{G}(h) \) of the Gaussian distribution \( G \) are also Gaussian with covariance operator \( \text{cov} \hat{G}(h) = \text{cov} G = b^2 I_d \). Therefore,
\[
\left| \hat{R}_h^{(s)}(t) \right| \leq \exp \left( -b^2 \| t \|^2 / 2 \right), \quad t, h \in \mathbb{R}^d, \| h \| \tau < 1.
\]
(4.14)

Using (4.12)–(4.14), we get, for \( t, h \in \mathbb{R}^{2d}, \| h \| \tau < 1: \)
\[
\left| \hat{M}_h^{(s)}(t) \right| \leq \prod_{\mu=1}^{2} \exp \left( -b^2 \| t^{(\mu)} \|^2 / 2 \right) = \exp \left( -b^2 \| t \|^2 / 2 \right).
\]
(4.15)

Applying (2.40), (4.10) and (4.15) with \( t = A^* u \) and \( h = A^* \gamma \), we see that
\[
\left| \hat{Q}_\gamma^{(s)}(u) \right| \leq \exp \left( -b^2 \| A^* u \|^2 / 2 \right) \leq \exp \left( -b^2 \| u \|^2 \right),
\]
(4.16)

for \( u, \gamma \in \mathbb{R}^{2d}, \| \gamma \| \sqrt{2} \tau < 1 \). The relations (4.11) and (4.16) imply that
\[
\left| \hat{F}_h(t) \right| \leq \exp \left( -b^2 \| t \|^2 \cdot 2^{n-2} \right), \quad t, h \in \mathbb{R}^j, \| h \| \sqrt{2} \tau < 1.
\]
(4.17)

It is clear that it suffices to verify (4.17) for \( j = 2d \) (for \( 1 \leq j < 2d \) one should apply (4.17) for \( j = 2d \) and for \( t, h \in \mathbb{R}^{2d}, \) with \( h_m = t_m = 0, m = j + 1, \ldots, 2d \)).

Using (4.6), (4.7) and (4.17), we see that
\[
\int_{\mathbb{R}^j} \left| \hat{F}_h(t) \right| dt \leq \exp \left( -b^2 \cdot 2^{n-3} \right) \int_{\mathbb{R}^j} \exp \left( -b^2 \| t \|^2 \cdot 2^{n-3} \right) dt
\]
\[
= \frac{(2\pi)^{j/2}}{(b^2 \cdot 2^{n-2})^{j/2}} \exp \left( -b^2 \cdot 2^n \right) \frac{(2\pi)^{j/2}}{(b^2 \cdot 2^n)^{j/2}} \leq \frac{(2\pi)^{j/2}}{(\det D)^{1/2} \tau^{2j}} \leq \frac{(2\pi)^{j/2}}{(2^n/\tau)^{j/2} \tau^{2j}} \leq \frac{(2\pi)^{j/2}}{(2^n/\tau)^{1/2}},
\]
(4.18)

if \( c_1 \) is small enough. The relations (4.7) and (4.18) imply (4.5). It remains to apply Theorem 2.1 to complete the proof of Theorem 1.1. □
End of the proof of Theorem 1.2  Let now the distributions \( \mathcal{L}(\xi_s) \) satisfy the conditions of Theorem 1.2. In this case, according to (2.8) and (4.2), we have
\[
\left| \hat{R}_h^{(s)}(t) \right| = \frac{\left( 4 + \tau^2(d + \|h\|^2 + 2i\langle h, t \rangle - \|t\|^2) \right) \exp \left( (\|h\|^2 + 2i\langle h, t \rangle - \|t\|^2)/2 \right)}{\left( 4 + \tau^2(d + \|h\|^2) \right) \exp \left( \|h\|^2/2 \right)}
\leq \left( 2 + \|t\|^2 \right) \exp \left( -\|t\|^2/2 \right)
\leq c_{25} \exp \left( -\|t\|^2/4 \right), \quad \|h\| \tau < 1. \tag{4.19}
\]
The rest of the proof is omitted. It is similar to that of Theorem 1.1 with \( b_2 = 1/2 \). The presence of \( c_{25} \) in the right-hand side of (4.19) can be easily compensated by choosing \( c_1 \) to be sufficiently small.

End of the proof of Theorem 1.3  Consider the dyadic scheme with
\[
\mathcal{L}(\xi_s) = \mathcal{L}(X_s) = L^{(s)} P, \quad s = 1, \ldots, 2^N. \tag{4.20}
\]
Putting \( H := \mathcal{L}(\zeta), \quad \psi_h(x) = e^{i\langle h, x \rangle} p(x), \quad h, x \in \mathbb{R}^d, \) and integrating by parts, we see that (for \( t \in \mathbb{R}^d, \ t \neq 0 \))
\[
\hat{H}_h(t) = \left( \mathbf{E} e^{i\langle h, \zeta \rangle} \right)^{-1} \int_{\|x\| \leq b_1} e^{i\langle t, x \rangle} \psi_h(x) \, dx
\]
\[
= -\left( \mathbf{E} e^{i\langle h, \zeta \rangle} \right)^{-1} \int_{\|x\| \leq b_1} e^{i\langle t, x \rangle} i \|t\|^2 d_t \psi_h(x) \, dx, \tag{4.21}
\]
where \( H_h = \overline{\Pi}(h) \). Besides, using (1.9), we see that
\[
\sup_{\|x\| \leq b_1} \sup_{\|h\| \leq b_2} \left| d_t \psi_h(x) \right| \leq b_5 \|t\|. \tag{4.22}
\]
As in the formulation of Theorem 1.3 we denote by \( b_m \) different positive quantities depending on \( H \). Note that the quantities depending on the dimension \( d \) can be considered as depending on \( H \) only as well. From (4.21) and (4.22) it follows that
\[
\sup_{\|h\| \leq b_1} \left| \hat{H}_h(t) \right| \leq b_6 \|t\|^{-1} \tag{4.23}
\]
(note that, by the Jensen inequality, \( \mathbf{E} e^{i\langle h, \zeta \rangle} \geq e^{\mathbf{E} \langle h, \zeta \rangle} = 1 \)). The inequality (4.23) implies that
\[
\sup_{\|h\| \leq b_1} \left| \hat{H}_h(t) \right| \leq \left( 1 + \frac{\|t\|}{b_7} \right)^{-1} \quad \text{for} \quad \|t\| \geq b_7 = 2b_6 \tag{4.24}
\]
and
\[
\sup_{\|h\| \leq b_1} \sup_{\|t\| \geq b_7} \left| \hat{H}_h(t) \right| \leq \frac{1}{2}. \tag{4.25}
\]
Since the distributions \( H_h \) are absolutely continuous, the relation \( |\hat{H}_h(t)| = 1 \) can be valid for \( t = 0 \) only. Furthermore, the function \( |\hat{H}_h(t)| \) considered as a function of two variables \( h \) and \( t \) is continuous for all \( h, t \in \mathbb{R}^d \). Therefore,

\[
\sup_{\|h\|_{b_2} \leq 1} \sup_{b_s \leq \|t\| \leq b_7} |\hat{H}_h(t)| \leq b_9 < 1, \tag{4.26}
\]

where

\[
b_8 = (4\sqrt{2} b_2 d)^{-1} \quad \text{and} \quad b_9 \geq 1/2. \tag{4.27}
\]

The inequalities (4.25) and (4.26) imply that

\[
\sup_{\|h\|_{b_2} \leq 1} \sup_{\|t\| \geq b_8} |\hat{H}_h(t)| \leq b_9 := e^{-b_{10}} < 1. \tag{4.28}
\]

Denoting \( L_h^{(s)} = \mathcal{L}^{(s)}(h) \), \( h \in \mathbb{R}^d \), \( s = 1, \ldots, 2^N \), and using (1.11), (1.12), (2.3) and (2.8), it is easy to see that

\[
\hat{H}_h^{(s)}(t) = \left( \hat{H}_{h, \sqrt{m}}(t/\sqrt{m}) \right)^{m} \hat{L}_h^{(s)}(t). \tag{4.29}
\]

The relations (1.10), (4.24), (4.28) and (4.29) imply that

\[
\sup_{\|h\| \tau \leq 1} \left| \hat{R}_h^{(s)}(t) \right| \leq \left( 1 + \frac{\|t\|}{b_7 \sqrt{m}} \right)^{-m} \quad \text{for} \quad \|t\| \geq b_7 \sqrt{m}, \tag{4.30}
\]

and

\[
\sup_{\|h\| \tau \leq 1} \sup_{\|t\| \geq b_8 \sqrt{m}} \left| \hat{R}_h^{(s)}(t) \right| \leq e^{-m b_{10}}. \tag{4.31}
\]

Using (4.12), (4.13), (4.20) and (4.30), we get, for \( r = 2^{n-1} \cdot 2k + 1, \ldots, 2^{n-1}(2k + 1) \), \( \|t\| \geq b_7 \sqrt{2m} \), \( t \in \mathbb{R}^{2d} \),

\[
\sup_{\|h\| \tau \leq 1} \left| \hat{M}_h^{(r)}(t) \right| \leq \min_{\mu = 1, 2} \left( 1 + \frac{\|t^{(\mu)}\|}{b_7 \sqrt{m}} \right)^{-m} \leq \left( 1 + \frac{\|t\|}{b_7 \sqrt{2m}} \right)^{-m}. \tag{4.32}
\]

Moreover,

\[
\sup_{\|h\| \tau \leq 1} \sup_{\|t\| \geq b_8 \sqrt{2m}} \left| \hat{M}_h^{(r)}(t) \right| \leq e^{-m b_{10}}. \tag{4.33}
\]

Using (2.40), (4.10), (4.32) and (4.33), we see that, for the same \( r \) and for \( t \in \mathbb{R}^{2d} \), \( \|t\| \geq b_7 \sqrt{m} \),

\[
\sup_{\|h\| \sqrt{2} \tau \leq 1} \left| \hat{Q}_h^{(r)}(t) \right| \leq \left( 1 + \frac{\|t\|}{b_7 \sqrt{m}} \right)^{-m} \tag{4.34}
\]

and

\[
\sup_{\|h\| \sqrt{2} \tau \leq 1} \sup_{\|t\| \geq b_8 \sqrt{m}} \left| \hat{Q}_h^{(r)}(t) \right| \leq e^{-m b_{10}}. \tag{4.35}
\]

It is easy to see that the relations (4.11), (4.34) and (4.35) imply that, for \( h \in \mathbb{R}^j \), \( \|h\| \sqrt{2} \tau < 1 \), and for \( t \in \mathbb{R}^j \), \( \|t\| \geq b_7 \sqrt{m} \),

\[
|\hat{F}_h(t)| \leq \left( 1 + \frac{\|t\|}{b_7 \sqrt{m}} \right)^{-m 2^{n-1}}. \tag{4.36}
\]
and
\[
\sup_{\|t\| \geq b_8 \sqrt{m}} \left| \hat{F}_h(t) \right| \leq e^{-mb_{10} \cdot 2^{n-2}}. \tag{4.37}
\]
It suffices to prove (4.36) and (4.37) for \( j = 2d \) (for \( 1 \leq j < 2d \) one should apply (4.36) and (4.37) for \( j = 2d \) and for \( h \in \mathbb{R}^{2d}, \|h\| \sqrt{2} \tau < 1, \ t \in \mathbb{R}^{2d} \) with \( h_m = t_m = 0, m = j + 1, \ldots, 2d \).

Note now that the set \( T \) defined in (4.6) satisfies the relation
\[
T \subset \left\{ t \in \mathbb{R}^j : \|t\| \geq b_8 \sqrt{m} \right\} \tag{4.38}
\]
(see (1.10) and (4.27)). Below (in the proof of (4.5)) we assume that \( \|h\| \sqrt{2} \tau < 1 \). According to (4.37) and (4.38), for \( t \in T \) we have
\[
\left| \hat{F}_h(t) \right|^{1/2} \leq e^{-mb_{10} \cdot 2^{n-2}}. \tag{4.39}
\]
Taking into account that \( \left| \hat{F}_h(t) \right| \leq 1 \), and \( m \geq b_4 \), choosing \( b_4 \) to be sufficiently large and using (1.10), (4.7), (4.36) and (4.39), we obtain
\[
\int_T \left| \hat{F}_h(t) \right| dt \leq \exp \left( -mb_{10} \cdot 2^{n-2} \right) \left( \int_{\mathbb{R}^j} \left( 1 + \frac{\|t\|}{b_7 \sqrt{m}} \right)^{-m \cdot 2^{n-2}} dt + b_{11} m^{d/2} \right)
\leq b_{12} m^{d/2} \exp \left( -mb_{10} \cdot 2^{n-2} \right)
\leq \frac{(2\pi)^{j/2} \sqrt{2} b_2 j^{3/2}}{m^{1/2} \cdot 2^{n/2} \cdot 2^{n_3/2}} = \frac{(2\pi)^{j/2} \sqrt{2} \tau^{3/2}}{\sigma \left( \det D \right)^{1/2}}. \tag{4.40}
\]
The inequality (4.5) follows from (4.40) immediately. It remains to apply Theorem 2.1. □

**Proof of Theorem 1.4** Define \( m_0, m_1, m_2, \ldots \) and \( n_1, n_2, \ldots \) by
\[
m_0 = 0, \quad m_s = 2^s, \quad n_s = m_s - m_{s-1}, \quad s = 1, 2, \ldots. \tag{4.41}
\]
It is easy to see that
\[
\log n_s \leq \log m_s = 2^s \log 2, \quad s = 1, 2, \ldots. \tag{4.42}
\]
By Assertion A (see (1.5)), for any \( s = 1, 2, \ldots \) one can construct on a probability space a sequence of i.i.d. \( X_1^{(s)}, \ldots, X_{n_s}^{(s)} \) and a sequence of i.i.d. Gaussian \( Y_1^{(s)}, \ldots, Y_{n_s}^{(s)} \) so that \( \mathcal{L}(X_k^{(s)}) = \mathcal{L}(\xi), \ \mathbb{E} Y_k^{(s)} = 0, \ \text{cov} Y_k^{(s)} = I_d \), and
\[
P \left\{ c_2 \Delta_s \geq \tau d^{3/2} \left( c_3 \log^4 d \log n_s + x \right) \right\} \leq e^{-x}, \quad x \geq 0, \tag{4.43}
\]
where
\[
\Delta_s = \max_{1 \leq r \leq n_s} \left| \sum_{k=1}^r X_{k}^{(s)} - \sum_{k=1}^r Y_{k}^{(s)} \right|. \tag{4.44}
\]
It is clear that we can define all the vectors mentioned above on the same probability space so that the collections $\Xi_s = \{X_1^{(s)}, \ldots, X_{n_s}^{(s)}, Y_1^{(s)}, \ldots, Y_{n_s}^{(s)}\}, \ s = 1, 2, \ldots$ are jointly independent. Then we define $X_1, X_2, \ldots$ and $Y_1, Y_2, \ldots$ by

$$X_{m_s-1+k} = X_k^{(s)}, \quad k = 1, \ldots, n_s, \ s = 1, 2, \ldots$$

(4.45)

In order to show that these sequences satisfy the assertion of Theorem 1.4, it remains to verify the equality (1.13).

Put

$$c_{25} = \frac{(c_3 \log 2 + 1)}{c_2}, \quad c_{26} = c_{25} \sum_{l=0}^{\infty} 2^{-l/2} = \frac{c_{25} \sqrt{2}}{\sqrt{2} - 1},$$

(4.46)

and introduce the events

$$A_l = \{ \omega : \Delta^{(l)} \geq 2^l c_{26} \tau d^{3/2} \log^* d \}, \quad l = 1, 2, \ldots$$

(4.47)

where

$$\Delta^{(l)} = \max_{1 \leq r \leq m_l} \left| \sum_{j=1}^{r} X_j - \sum_{j=1}^{r} Y_j \right|.$$ 

(4.48)

According to (4.44), (4.45) and (4.48), we have

$$\Delta^{(l)} \leq \Delta_1 + \cdots + \Delta_l.$$ 

(4.49)

Taking into account the relations (4.42), (4.46), (4.47), (4.49) and applying the inequality (4.43) with $x = 2^{(s+l)/2}$, we get

$$\mathbb{P}\{A_l\} \leq \sum_{s=1}^{l} \mathbb{P}\{\Delta_s \geq 2^{(s+l)/2} c_{25} \tau d^{3/2} \log^* d\} \leq \sum_{s=1}^{l} \exp \left( - 2^{(s+l)/2} \right) \leq c \exp \left( - 2^{l/2} \right).$$

(4.50)

The inequality (4.50) implies that $\sum_{l=1}^{\infty} \mathbb{P}\{A_l\} < \infty$, hence, by the Borel–Cantelli lemma with probability one a finite number of the events $A_l$ occurs only. This implies the equality (1.13) with $c_4 = 2 c_{26}/ \log 2$ (see (4.41), (4.47) and (4.48)). □

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