Psychometric properties of Brief-COPE inventory among nurses

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Abstract

Objective: Coping with persistent demands in modern healthcare and facing new challenges such as the COVID-19 pandemic, is increasingly necessary among nurses. Brief-COPE inventory is widely used to assess coping, however, validation evidence is absent and previous results were inconsistent. This study aimed to validate psychometric properties of Brief-COPE inventory to ensure culturally appropriate usage.

Methods: Preliminary validation study on 423 female nurses from the United Arab Emirates. Confirmatory factor analysis (CFA) with maximum likelihood estimation was computed to test four different previous models. Exploratory factor analysis (EFA) protocol was used to determine underlying domain structure of Brief-COPE in this population.

Results: The four previous models on CFA had inadequate fit indexes. Two-factor (22-items) second-order model that explained 37.0% of the total variance with Cronbach's alpha at 0.81 and 0.88, respectively, was suggested.

Conclusion: This validation of Brief-COPE is timely considering nurses enduring different types of stressors in addition to the COVID-19 pandemic situation. In addition, cultural diversity need to be considered in coping research. Re-assessment of this exploratory structure is necessary to ensure accurate measurement of coping strategies.

Highlights

- Nurses faces persistent demands and new challenges in midst of COVID-19
- Coping strategies are important mechanism to mitigate adverse effects of high stress environment that has personal and organisational implications
- Brief-COPE has been extensively used in nursing research, however, its psychometric properties need to be evaluated
- Cultural diversity has significant influence towards the constructs of brief-COPE inventory

Introduction

Coping with persistent demands and challenges in the modern healthcare system is increasingly necessary among healthcare professionals (Robertson et al., 2016). Nurses, in particular, faces high stress and burnout levels from increased workload and work engagement, emotional exhaustion, staff shortages and poor health (Yu, Raphael, Mackay, Smith, & King, 2019). Studies have utilized different measurement tools to assess nurses’ coping strategies, in various settings (Isa et al., 2019; Lee, Kuo, Chien, & Wang, 2016). One of the most widely used measures of nurses’ coping strategies is the Brief-COPE (Coping Orientation to Problems Experienced) inventory, a shortened version of the full 60-items (16 scales) COPE inventory developed by Carver, Scheier, & Weintraub (1989). Based on the Folkman & Lazarus’ Ways of Coping model and the Behavioural self-regulation model, the inventory aimed to assess various positive and negative coping strategies effectively (Carver et al., 1989). The inventory was further refined to reduce participant fatigue and item redundancies and eventually the creation of the Brief-COPE inventory, which consists of 28-items with 14 scales (Carver, 1997).

Following sound theoretical guidance, a number of health-relevant studies have provided empirical evidences by using statistical tools to extract underlying domains and test psychometric properties in various settings and populations such as caregivers of people with dementia in the United Kingdom (Cooper, Katona, & Livingston, 2008), among breast cancer chemotherapy patients in Malaysia (Yusoff, Low, & Yip, 2009), people with traumatic brain injury in New Zealand (Snell, Siegert, Hay-Smith, & Surgenor, 2011), people living with HIV in China (Su et al., 2015), person living with HIV/AIDS in India (Mohanraj et al., 2015), pregnant minority women in the United States (Rui et al., 2015), adults in Italy (Monzani et al., 2015), and community population in Chile (García, Barraza-Peña, Wlodarczyk, Alvear-Carrasco, & Reyes-Reyes, 2018). Brief-COPE is also translated in a number of languages such as French version (Muller & Spitz, 2003), Malay version (Yusoff et al., 2009), Brazilian-Portuguese version (Brasileiro et al., 2016), and Chilean version (Garcia et al., 2018).
Although several coping strategies instruments exists such as the Folkman and Lazarus’s Ways of Coping questionnaire (Folkman & Lazarus, 1988), the Multidimensional Coping Inventory (Endler & Parker, 1990) and the Coping Inventory for stressful situation (Endler, Kantor, & Parker, 1994), the relatively lengthy items were not particularly useful in long research protocols and clinical research. The 28-items Brief-COPE do not have this shortcoming, and have been used in many studies extensively. However, the validation results have not always been consistent due to the complex nature of coping dimensions on different types of stressors and study populations. In addition, we have not found studies examining validity and reliability estimates of Brief-COPE on nurses in the United Arab Emirates (UAE). Therefore, this paper aimed to bridge this gap by validating the psychometric properties of Brief-COPE inventory to ensure culturally appropriate usage.

**Materials And Methods**

**Study design and setting**

This preliminary validation study investigated the psychometric properties of the English version of Brief-COPE inventory among nurses working in the United Arab Emirates (UAE). Participants were recruited between October 2017 and December 2018 using cluster sampling.

**Eligible population**

Eligible participants were female nurses aged \( \geq 20 \) years who were not pregnant. We excluded men in the current study because the number of male nurses is very minimal in the country. Participants who had chronic health conditions (e.g., diabetes mellitus, hypertension, coronary artery disease, renal disease) and those taking oral contraceptive pills were excluded from this study.

**Data collection procedure**

Trained research assistants visited the selected clusters to recruit participants. Advertisements about the study including an invitation to participate were sent to the selected sites before the site visits. The research assistants then met nurses, screened interested nurses for eligibility, introduced the study, obtained consent, and asked participants to complete the questionnaire.

**Research instrument**

The original brief-COPE by Carver (1997) comprised of two-items in each 14 subscales including self-distraction, active coping, denial, substance use, use of emotional support, use of instrumental support, behavioral disengagement, venting, positive reframing, planning, humor, acceptance, religion, and self-blame. Meyer (2001) categorized these subscales into second-order factor model, which consisted of “Adaptive coping strategies” (use of emotional support, positive reframing, acceptance, religion, humor, active coping, planning, and use of instrumental support) and “Maladaptive coping strategies” (venting, denial, substance use, behavioural disengagement, self-distraction, and self-blame). Cooper, Katona, Orrell, & Livingston (2006) further categorized the original subscales into three, which consisted of “Emotion-focused strategies” (use of emotional support, positive reframing, acceptance, religion, and humor), “Problem-focused strategies” (active coping, planning, and use of instrumental support) and “Dysfunctional coping strategies” (venting, denial, substance use, behavioural disengagement, self-distraction, and self-blame). More recently, Eisenberg, Shen, Schwarz, & Mallon (2012) indicated two major, i.e., “Approach coping” (active coping, emotional support, use of instrumental support, positive reframing, planning, and acceptance) and “Avoidant coping” (self-distraction, denial, substance use, behavioural disengagement, venting, and self-blame) that excludes humor and religion.

**Statistical analyses**

Brief-COPE inventory has 4-point Likert response from 1 (I haven't been doing this at all) to 4 (I've been doing this a lot). Mean and standard deviation of the item scores and observed range as well as floor and ceiling effects where more than 15% is considered significant effect (Terwee et al., 2007), were calculated. To evaluate the construct validity of the brief-COPE inventory, a series of confirmatory factor analysis (CFA) with maximum likelihood estimation was computed to test four different previous models including the original structure by Carver (1997), and other studies by Meyer (2001), Cooper et al. (2006), and Eisenberg et al. (2012) that are labelled as Model 1, Model 2, Model 3, and Model 4 respectively. Various model fit
indices were considered: 1) relative chi-square (χ²), 2) χ²/df where less than 2 is good fit, 3) root mean square error of approximation (RMSEA) where ≤ 0.08 is considered good fit, 4) Non-normed fit index (NFI) where ≥ 0.95 is considered good fit, 5) Comparative fit index (CFI) where ≥ 0.90 is considered good fit, 6) Goodness-of-fit index (GFI) where ≥ 0.95 is considered good fit, and 7) Akaike's information criterion (AIC) where lower values is considered better model (Hu & Bentler, 1999).

The four models had inadequate fit indexes; thus we performed exploratory factor analysis (EFA) model using maximum likelihood technique with varimax rotation to determine number of factors extracted and ascertained with scree plot, eigenvalue value more than 1, and factor loadings of more than 0.4. Prior to this, KMO value (≥ 0.5) and Bartlett's sphericity test (p < 0.001) were computed (Dziuban & Shirkey, 1974). Finally, a structural equation diagram with standardized regression estimates was rendered to visualize the EFA model. Item-total correlation coefficients and Cronbach's alpha coefficients where ≥ 0.70 is considered good internal consistency (Streinmer, 1993) were also estimated. All analysis was computed using R 3.6.3 and RStudio for Mac.

**Ethical considerations**

This study was approved by [masked]. Participants were informed that all personal information will be kept confidential and saliva samples and questionnaires will be give serial number without personal identifications. Written consent forms were obtained from participants and kept in a locked cabinet by the research team.

**Results**

423 valid data from nurses in UAE who completed the brief-COPE questionnaire, were included in this analysis. Participants had mean age of 36.7 years (SD 8.5 years). Majority was expatriates (85.1%) from different nationalities and 68.4% were married.

Table 1 summarize the results of CFA model fit indices with respect to Model 1 to Model 4. Chi-square test of absolute model fit (p < 0.001) and other indices including NFT, CFI and GFI indicated all models had poor absolute fit. RMSEA index indicated only Model 1 had cut-off ≤ 0.08, a good fit. AIC considered Model 4 was comparatively a better model than the others.

| Scales | Dimension | Items | χ² | χ²/ df | RMSEA | NFI | CFI | GFI | AIC  |
|--------|-----------|-------|----|--------|--------|-----|-----|-----|------|
| Model 1 | 14        | 0     | 28 | 940.0  | 3.63   | 0.079 | 0.802 | 0.845 | 0.866 | 26950 |
| Model 2 | 14        | 2     | 28 | 1485.7 | 4.48   | 0.091 | 0.688 | 0.737 | 0.793 | 27350 |
| Model 3 | 14        | 3     | 28 | 1511.8 | 4.54   | 0.091 | 0.682 | 0.731 | 0.791 | 27374 |
| Model 4 | 12        | 2     | 24 | 1197.6 | 5.01   | 0.097 | 0.679 | 0.722 | 0.805 | 23507 |

Model 1 (Carver, 1997) Model 3 (Cooper et al., 2006)
Model 2 (Meyer, 2001) Model 4 (Eisenberg et al., 2012)
χ²(df) (Chi-square (degree of freedom)) RMSEA (Root Mean Square of approximation)
NFI (Non-normed fit index) SRMR (Standardized Root Mean Square Residual)
CFI (Comparative fit index) GFI (Goodness of fit index)
AGFI (Adjusted Goodness of fit index) AIC (Akaike information criterion)
Based on the original 14 subscales of brief-COPE inventory, CFA resulted on highly cross-loaded estimates on all factors except for behavioural disengagement, self-blame, substance use, humor, religion, and acceptance. Emotional support and use of informational support loaded on the same factor. Likewise, positive reframing and planning also formed a single factor.

Due to poor model fit with CFA models and high cross-loadings, EFA was conducted to establish factor structure for our study population. The KMO value (0.85) and Bartlett's sphericity test ($\chi^2 = 4632.7$, df = 378, $p < 0.001$) indicated good factorability. The scree plot and eigenvalue more than 1 suggested seven factors with five point optimal extraction (Fig. 1). Varimax rotation with factor loadings more than 0.4 yielded two factor extraction (11 scales excluded humor, self-distraction, and substance use), derives EFA model accounting for 37% of the total variance with the following fit indices: $\chi^2$/df = 4.74, RMSEA = 0.099, NFI = 0.75, GFI = 0.84, AIC = 21790.8. Item-total correlation coefficients ranged from 0.37 to 0.69 and 0.52 to 0.63 for Factor 1 and Factor 2, respectively. Table 2 details the loadings and reliability estimates for this EFA model as well as mean and standard deviation of the scores and floor and ceiling effects.
Table 2
Mean scores, Distribution, Factor loadings and Reliability estimates of Brief-COPE (n = 423)

| Scales                  | Item no. | Item no. | Mean score (SD) | Floor (%) | Ceiling (%) | Factor Loadings | Alpha (scales) | Alpha (factor) |
|-------------------------|----------|----------|-----------------|-----------|-------------|-----------------|----------------|----------------|
| **Factor 1** Self-distraction | 1        |          | 5.1 (1.5)       | 5.0       | 5.0         | -               | 0.81           |                |
|                         | 19       |          |                 |           |             |                 |                |                |
| Denial                  | 3        | 8        | 3.7 (1.5)       | 28.8      | 1.9         | 0.45            | 0.56           |                |
|                         |          |          |                 |           |             | 0.63            |                |                |
| Substance use           | 4        | 11       | 2.4 (0.9)       | 81.3      | 0.2         | -               | -              |                |
|                         |          |          |                 |           |             |                 |                |                |
| Behavioral disengagement| 6        | 16       | 4.2 (1.7)       | 19.1      | 5.4         | 0.58            | 0.68           | 0.65           |
| Vventing                | 9        | 21       | 4.7 (1.5)       | 6.6       | 5.0         | 0.53            | 0.45           | 0.31           |
| Self-blame              | 13       | 26       | 4.5 (1.7)       | 14.4      | 5.7         | 0.73            | 0.75           | 0.73           |
| Humor                   | 18       | 28       | 4.5 (1.8)       | 19.6      | 5.9         | -               | -              |                |
|                         |          |          |                 |           |             |                 |                |                |
| **Factor 2** Active coping | 2        | 7        | 5.8 (1.4)       | 1.4       | 11.1        | 0.48            | 0.53           | 0.88           |
| Emotional support       | 5        | 15       | 5.3 (1.6)       | 2.4       | 13.0        | 0.52            | 0.67           | 0.62           |
| Information support     | 10       | 23       | 5.3 (1.6)       | 3.5       | 12.5        | 0.54            | 0.61           | 0.61           |
| Positive reframing      | 12       | 17       | 5.6 (1.5)       | 2.1       | 13.2        | 0.49            | 0.54           | 0.65           |
| Planning                | 14       | 25       | 5.7 (1.5)       | 3.5       | 12.3        | 0.51            | 0.60           | 0.53           |
| Acceptance              | 20       | 24       | 5.7 (1.5)       | 2.8       | 16.3        | 0.64            | 0.65           | 0.58           |
| Religion                | 22       | 27       | 6.4 (1.6)       | 1.7       | 35.7        | 0.61            | 0.79           |                |

Exploratory factor analysis (Varimax) SD = Standard deviation Alpha = Cronbach’s alpha

Initial analysis showed skewed distributions of the responses on several items. Floor effects were present for denial, substance use, and behavioural disengagement. Ceiling effects were significant for acceptance subscale. Although
Cronbach's alpha coefficient for first-order subscales were mostly below 0.70 except for self-blame and religion, there was good alpha estimate for the second-order two-factor EFA model at 0.81 and 0.86, respectively, where the structure is illustrated in Fig. 2.

**Discussion**

This study has established the psychometric properties of the Brief-COPE inventory among nurses working in the UAE. Similar to previous studies, our validation estimates have demonstrated inconsistencies with previous models due to the complex nature of coping especially in different parts of the world that could be influenced by culture, different generations, types of stressors, and study populations. This study is timely considering that COVID-19 pandemic has affected health professionals worldwide, and nurses as front liners, is even more pressing to ensure instruments measuring their coping strategies are both valid and reliable (Huang, xu, & Liu, 2020; Siyu et al., 2020; Zhuang et al., 2020).

A number of studies has used Brief-COPE to examine nurses coping strategies (Alharbi & Alshehry, 2019; Fathi & Simamora, 2019; McMeekin, Hickman Jr, Douglas, & Kelley, 2017), however, in the absence of validation study, the results reported may be problematic and disputed. In particular, the CFA models using structure from previous studies did not achieve good fit, observed high cross-loadings, and Cronbach's alpha of individual subscales were inadequate (although most were above 0.5), which was similarly reported in previous validation studies including the original Brief-COPE study (Charles S. Carver, 1997; García et al., 2018; Monzani et al., 2015; Su et al., 2015). Nevertheless, the developer of Brief-COPE reported that the flexibility of the inventory encourage researchers to apply and test on different settings and population to extensively either using all the subscales or selectively use relevant scales in their research (Charles S. Carver, 1997).

Even though Brief-COPE was developed with sound theoretical model, existing validation studies of different version of brief-COPE have produced large variation in empirical evidences, particularly on the dimensionality of the scales, which may limit comparability of results in studies using this inventory. Mohanraj et al. (2015) exploration of the Tamil version of Brief-COPE underlying structure among person living with HIV/AIDS in India, yielded a five factor (17-items) model accounting 41.5% of the total variance. Yusoff et al. (2009) investigation on Malay version of Brief-COPE among adolescent in Malaysia, produced a nine factor model explaining 71.2% of the total variance. Su et al. (2015) generated a six factor model on the Chinese version of Brief-COPE, which explained 55.5% of total variance. Reich, Costa-Ball, & Remor (2016) constructed a four factor (24-items) model of the Uruguayan version, and more recently, Gloria & Peña (2018) produced an eight-factor (24-items) model on the Chilean version. In our study examining nurses in the UAE, we suggested a two-factor (22-items) model that explained 37.0% of the total variance. Substance use subscale was excluded because UAE is an Islamic country, and items on alcohol is not relevant. Self-distraction and Humor subscales were slightly difficult to justify, however, considering highest score was Religion subscale, we could postulate that religiosity would most likely have preceded them.

**Limitations**

We expected that the reduced items will compromise theoretical content of the inventory and future researcher should attempt to re-establish content validity with area experts and replicate the results of our study, and if possible, to conduct full validation study incorporating estimates for criterion validity and test-retest reliability.

**Conclusion**

Overall, a two-factor second-order model provides a good model to interpret coping strategies among nurses in the UAE, which is timely considering nurses around the world is currently battling different types of stressors in addition to the COVID-19 pandemic situation. Our study also highlighted the need to account for cultural diversity in coping research as studies on translated version of Brief-COPE had yielded inconsistent underlying constructs. Re-assessment of this exploratory structure is critically warranted on nurses from different study settings and population to ensure accurate measurement of coping strategies.
Declarations

Ethics approval and consent to participate

This study was approved by the Research Ethics Committee of the Ministry of Health and Prevention, UAE (MOHAP/DXB/SUBC/No. 6/2017) and by the Research Ethics Committee of Sharjah Institute for Medical and Health Sciences Research at the study institution (REC-16-09-07). Written consent was obtained from participants prior to study.

Consent for publication

All authors approved the manuscript for publication.

Availability of data and materials

No data will be deposited.

Competing interest

No conflict of interest has been declared by the authors.

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Authors’ contributions

HAR concept design, data analysis, writing manuscript and final approval.

WBI study design, data collection, critical review and final approval.

LN concept design, reviewed data analysis, critical review and final approval.

References

Alharbi, H., & Alshehry, A. (2019). Perceived stress and coping strategies among ICU nurses in government tertiary hospitals in Saudi Arabia: a cross-sectional study. *Annals of Saudi Medicine, 39*(1), 48–55.

Brasileiro, S. V., Orsini, M. R. C. A., Cavalcante, J. A., Bartholomeu, D., Montiel, J. M., Costa, P. S. S., & Costa, L. R. (2016). Controversies regarding the psychometric properties of the Brief COPE: The case of the Brazilian-Portuguese version “COPE Breve.” *PLoS ONE, 11*(3), 1–13. https://doi.org/10.1371/journal.pone.0152233

Carver, C., Scheier, M., & Weintraub, J. (1989). Assessing Coping Strategies: A Theoretically Based Approach, Journal of Personality and Social Psychology, *56*(2), 267–283.

Charles S. Carver. (1997). You want to measure coping but your protocol’s too long: Consider the brief COPE. *International Journal of Behavioral Medicine, 4*(1), 92–100.

Cooper, C., Katona, C., & Livingston, G. (2008). Validity and Reliability of the Brief COPE in Carers of People With Dementia: The LASER-AD Study. *The Journal of Nervous and Mental Disease, 196*(11). Retrieved from https://journals.lww.com/jonmd/Fulltext/2008/11000/Validity_and_Reliability_of_the_Brief_COPE_in.7.aspx

Cooper, C., Katona, C., Orrell, M., & Livingston, G. (2006). Coping strategies and anxiety in caregivers of people with Alzheimer's disease: The LASER-AD study. *Journal of Affective Disorders, 90*(1), 15–20. https://doi.org/https://doi.org/10.1016/j.jad.2005.08.017
Dziuban, C. D., & Shirkey, E. C. (1974). When is a correlation matrix appropriate for factor analysis? Some decision rules. *Psychological Bulletin, 81*(6), 358–361. https://doi.org/10.1037/h0036316

Eisenberg, S. A., Shen, B.-J., Schwarz, E. R., & Mallon, S. (2012). Avoidant coping moderates the association between anxiety and patient-rated physical functioning in heart failure patients. *Journal of Behavioral Medicine, 35*(3), 253–261. https://doi.org/10.1007/s10865-011-9358-0

Endler, N. S., & Parker, J. D. A. (1994). State-trait coping, state-trait anxiety and academic performance. *Personality and Individual Differences, 16*(5), 663–670. https://doi.org/10.1016/0191-8869(94)90208-9

Endler, N. S., & Parker, J. D. (1990). Multidimensional assessment of coping: A critical evaluation. *Journal of Personality and Social Psychology. US: American Psychological Association.* https://doi.org/10.1037/0022-3514.58.5.844

Fathi, A., & Simamora, R. H. (2019). Investigating nurses’ coping strategies in their workplace as an indicator of quality of nurses’ life in Indonesia: a preliminary study. In *IOP conference series: Earth and Environmental science* (Vol. 248, p. 12031). IOP Publishing.

Folkman, S., & Lazarus, R. S. (1988). Coping as a mediator of emotion. *Journal of Personality and Social Psychology. US: American Psychological Association.* https://doi.org/10.1037/0022-3514.54.3.466

García, F. E., Barraza-Peña, C. G., Wlodarczyk, A., Alvear-Carrasco, M., & Reyes-Reyes, A. (2018). Psychometric properties of the Brief-COPE for the evaluation of coping strategies in the Chilean population. *Psicologia: Reflexao e Critica, 31*(1). https://doi.org/10.1186/s41155-018-0102-3

Gloria, C., & Peña, B. (2018). Propiedades psicométricas del Cuestionario Brief Cope 28 en población Chilena expuesta a eventos altamente estresantes. In *XVI Coloquio Panamericano de Investigación en Enfermería.*

Hu, L., & Bentler, P. M. (1999). Cutoff criteria for fit indexes in covariance structure analysis: Conventional criteria versus new alternatives. *Structural Equation Modeling: A Multidisciplinary Journal, 6*(1), 1–55. https://doi.org/10.1080/10705519909540118

Huang, L., Xu, F. Ming, & Liu, H. Rong. (2020). Emotional responses and coping strategies of nurses and nursing college students during COVID-19 outbreak. *MedRxiv, 2020.03.05.20031898.* https://doi.org/10.1101/2020.03.05.20031898

Isa, K. Q., Ibrahim, M. A., Abdul-Manan, H.-H., Mohd-Salleh, Z.-A. H., Abdul-Mumin, K. H., & Rahman, H. A. (2019). Strategies used to cope with stress by emergency and critical care nurses. *British Journal of Nursing, 28*(1), 38–42. https://doi.org/10.12968/bjon.2019.28.1.38

Lee, H. F., Kuo, C. C., Chien, T. W., & Wang, Y. R. (2016). A Meta-Analysis of the Effects of Coping Strategies on Reducing Nurse Burnout. *Applied Nursing Research, 31,* 100–110. https://doi.org/10.1016/j.apnr.2016.01.001

McMeekin, D. E., Hickman Jr, R. L., Douglas, S. L., & Kelley, C. G. (2017). Stress and coping of critical care nurses after unsuccessful cardiopulmonary resuscitation. *American Journal of Critical Care, 26*(2), 128–135.

Meyer, B. (2001). Coping with Severe Mental Illness: Relations of the Brief COPE with Symptoms, Functioning, and Well-Being. *Journal of Psychopathology and Behavioral Assessment, 23*(4), 265–277. https://doi.org/10.1023/A:1012731520781

Mohanraj, R., Jeyaseelan, V., Kumar, S., Mani, T., Rao, D., Murray, K. R., & Manhart, L. E. (2015). Cultural adaptation of the Brief COPE for persons living with HIV/AIDS in southern India. *AIDS and Behavior, 19*(2), 341–351. https://doi.org/10.1007/s10461-014-0872-2

Monzani, D., Steca, P., Greco, A., D’Addario, M., Cappelletti, E., & Pancani, L. (2015). The Situational Version of the Brief COPE: Dimensionality and Relationships With Goal-Related Variables. *Europe's Journal of Psychology, 11*(2), 295–310.
Muller, L., & Spitz, E. (2003). [Multidimensional assessment of coping: validation of the Brief COPE among French population]. *L'Encephale, 29*(6), 507–518. Retrieved from http://europepmc.org/abstract/MED/15029085

Reich, M., Costa-Ball, C. D., & Remor, E. (2016). Psychometric Properties of the Brief COPE in a Sample of Uruguayan Women. *Avances En Psicología Latinoamericana, 34*(3), 615–636.

Robertson, H. D., Elliott, A. M., Burton, C., Iversen, L., Murchie, P., Porteous, T., & Matheson, C. (2016). Resilience of primary healthcare professionals: a systematic review. *British Journal of General Practice, 66*(647), e423 LP-e433. https://doi.org/10.3399/bjgp16X685261

Ruiz, R. J., Gennaro, S., O’Connor, C., Marti, C. N., Lulloff, A., Keshinover, T., … Melnyk, B. (2015). Measuring Coping in Pregnant Minority Women. *Western Journal of Nursing Research, 37*(2), 257–275. https://doi.org/10.1177/0193945914527176

Siyu, C., Xia, M., Wen, W., Cui, L., Yang, W., Liu, S., … Lei, W. (2020). Mental health status and coping strategy of medical workers in China during The COVID-19 outbreak. *MedRxiv, 2020.02.23.20026872*. https://doi.org/10.1101/2020.02.23.20026872

Snell, D. L., Siegert, R. J., Hay-Smith, E. J. C., & Surgenor, L. J. (2011). Factor Structure of the Brief COPE in People With Mild Traumatic Brain Injury. *The Journal of Head Trauma Rehabilitation, 26*(6). Retrieved from https://journals.lww.com/headtraumarehab/Fulltext/2011/11000/Factor_Structure_of_the_Brief_COPE_in_People_With.3.aspx

Streinmer, D. L. (1993). A Checklist for Evaluating the Usefulness of Rating Scales. *The Canadian Journal of Psychiatry, 38*(2), 140–148. https://doi.org/10.1177/07067437930380214

Su, X. you, Lau, J. T. F., Mak, W. W. S., Choi, K. C., Feng, T. jian, Chen, X., … Cheng, J. quan. (2015). A preliminary validation of the Brief COPE instrument for assessing coping strategies among people living with HIV in China. *Infectious Diseases of Poverty, 4*(1), 1–10. https://doi.org/10.1186/s40249-015-0074-9

Terwee, C. B., Bot, S. D. M., de Boer, M. R., van der Windt, D. A. W. M., Knol, D. L., Dekker, J., … de Vet, H. C. W. (2007). Quality criteria were proposed for measurement properties of health status questionnaires. *Journal of Clinical Epidemiology, 60*(1), 34–42. https://doi.org/10.1016/j.jclinepi.2006.03.012

Yu, F., Raphael, D., Mackay, L., Smith, M., & King, A. (2019). Personal and work-related factors associated with nurse resilience: A systematic review. *International Journal of Nursing Studies, 93*, 129–140. https://doi.org/10.1016/j.ijnurstu.2019.02.014

Yusoff, N., Low, W. Y., & Yip, C. H. (2009). Cope Scale: a Study on Malaysian Women Treated With. *MPJ Online Early, 1–9*. https://doi.org/10.15583/jpchrom.2017.003

Zhuang, Z., Zhao, S., Lin, Q., Cao, P., Lou, Y., Yang, L., & He, D. (2020). Preliminary estimation of the novel coronavirus disease (COVID-19) cases in Iran: a modelling analysis based on overseas cases and air travel data. *International Journal of Infectious Diseases*. 