Immune-related conditions and subsequent risk of brain cancer in a cohort of 4.5 million male US veterans

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Background: Case–control studies have reported an inverse association between self-reported history of allergy and risk of glioma, but cohort data are limited. Our objectives were to evaluate the associations of major groups of medically diagnosed immune-related conditions (allergy/atopy, autoimmune disease, diabetes, infectious/inflammatory disease) and to explore associations with specific conditions in relation to subsequent diagnosis of brain cancer in a large cohort study.

Methods: We used hospital discharge records for a cohort of 4.5 million male US veterans, of whom 4383 developed primary brain cancer. Rate ratios (RRs) and 95% confidence intervals (CIs) were calculated using time-dependent Poisson regression.

Results: We found a significant trend of decreasing RRs for brain cancer with longer duration of allergy/atopy ($P = 0.02$), but not for other conditions studied. Rate ratios of brain cancer for allergy/atopy and diabetes with duration of 10 or more years were 0.60 (95% CI: 0.43, 0.83) and 0.75 (95% CI: 0.62, 0.93), respectively. Several associations with specific conditions were found, but these did not withstand correction for multiple comparisons.

Conclusions: This study lends some support to an inverse association between allergy/atopy and diabetes of long duration and brain cancer risk, but prospective studies with biological samples are needed to uncover the underlying biological mechanisms.

Brain cancer is one of the deadliest types of malignancy, with 23 130 new cases and 14 080 deaths expected in the United States for 2013 (ACS, 2013). Previous studies have reported certain chemicals, occupations and dietary factors to be associated with brain cancer, but results are inconsistent. The most established risk factors include exposure to ionising radiation, familial cancer syndromes, and, most recently, several single-nucleotide polymorphisms identified through genome-wide association studies (Inskip et al, 1995; Bondy et al, 2008; Ohgaki, 2009; Rajaraman et al, 2012; Walsh et al, 2013). As the immune system is known to have a large role in oncogenesis, studying the effects of the dysregulated immune system in relation to primary brain cancer risk has also been of aetiological interest.

A number of case–control studies have reported an inverse relationship between history of allergy/atopy and risk of primary brain cancer (Ryan et al, 1992; Ciccittini et al, 1997; Chen et al, 2011; McCarthy et al, 2011), and two independent case–control studies have found an interaction between risk loci in the 9p21 region and history of allergy on risk of glioma, the most common type of primary brain cancer (Schoemaker et al, 2010; Lachance et al, 2011). A recent meta-analysis reported that risk of glioma associated with any self-reported allergic condition was significantly reduced; however, prospective cohort data contributed only 79 of the total 6408 cases (Chen et al, 2011).

Other immune-related conditions, such as autoimmune diseases, type II diabetes (T2DM), and infectious/inflammatory conditions, have not been extensively studied in relation to brain cancer risk, but may provide useful aetiological insights. Large cohort studies with clinically diagnosed immune conditions are needed to confirm reported associations for brain cancer, to establish new associations, and to clarify their dependence on time to diagnosis, age, and sex.
Our primary objective was to evaluate the associations of four major groups of clinically diagnosed immune-related conditions including allergy/atopy, autoimmune disease, diabetes and infectious/inflammatory disease in relation to subsequent risk of brain cancer using the hospital discharge records of a large cohort of 4.5 million black and white male US veterans. A key aspect of this design is that medical information was collected before, and independent of, the diagnosis of brain cancer. Our secondary objective was to explore relationships between specific, less common clinically diagnosed conditions and brain cancer risk.

MATERIALS AND METHODS

Study population and follow-up. This analysis was conducted using data from the Patient Treatment File of the US Veterans Administration (VA) medical system and was based on records of all inpatient hospitalisations at all 142 United States VA hospitals operating nationwide between 1 July 1969 and 30 September 1996. In total, there were 26 million hospital discharge records corresponding to 5,790,493 veterans with at least one hospital visit. Because of small percentages, veterans under 18 years or older than 100 years ($n=2,969; 0.05\%$), of female sex ($n=112,527; 1.9\%$), or with a race other than black or white ($n=135,651; 2.3\%$) were excluded from analysis (Brown et al, 2008). To study risk factors for incident cancers, patients who had cancer at first admission or who died on the day of first admission (666,650; 11.5\%) or developed cancer or died within the first year of first admission (371,129; 6.4\%) were also excluded. The final study population was comprised of 4,501,578 veterans and 4,383 incident, primary brain cancer cases.

Person-years at risk for brain cancer began 1 year after the first hospitalisation and ended at diagnosis of brain cancer, another primary malignancy, administrative censorship (30 September 1996), or death, whichever came first. Dates of death were ascertained through record linkage to Social Security Administration mortality files. The National Institutes of Health Office of Human Subjects Research granted exemption from institutional review board review and waived informed consent, because the study was restricted to existing data with all personal identifiers removed.

Outcome and exposures. Primary brain cancer cases were patients from the eligible population with an ICDA-8 (International Classification of Diseases, Adapted, 8th Revision) or ICD-9-CM (International Classification of Diseases, 9th Revision, Clinical Modification) discharge diagnosis of 191 (malignant neoplasm of the brain). This code primarily includes gliomas (e.g., astrocytoma, glioblastoma, ependymoma, oligodendroglioma, oligoastrocytoma) and some rare childhood tumours (e.g., medulloblastoma). Meningioma, acoustic neuroma, central nervous system lymphomas, and pituitary tumours are not classified under 191. Thus in our data set of middle-aged and older adults, we expect glioma to make up over 95\% of primary brain cancer cases (CBTRUS, 2012). Primary brain cancer diagnoses in the Patient Treatment File have been previously validated through medical record extraction in a smaller case–control study, with medical records matching 85\% of discharge diagnoses (Rollison et al, 2004).

Diagnoses of interest included allergic/atopic conditions, autoimmune diseases, diabetes, and infectious/inflammatory diseases that were obtained from hospital discharge records and have been previously categorised for this cohort (Koshiol et al, 2011). We used any discharge diagnosis code of the specified conditions irrespective of position in the file, and without making assumptions about the reason for admission. Specific allergic/atopic conditions included allergic rhinitis, asthma, eczema and dermatitis, erythema, and urticaria. Autoimmune diseases were categorised into those with detectable autoantibodies and systemic involvement (rheumatoid arthritis, systemic lupus erythematosus, systemic sclerosis, Sjogren’s syndrome), detectable autoantibodies and organ involvement (Addisons disease, amyotrophic lateral sclerosis, autoimmune haemolytic anaemia, Celiac disease, chronic rheumatic heart disease, discoid lupus erythematosus, Graves disease, Hashimoto thyroiditis, immune thrombocytopenic purpura, localised scleroderma, myasthenia gravis, pernicious anaemia, polyarteritis, primary biliary cirrhosis), and without detectable autoantibodies (amyotrophic lateral sclerosis, ankylosing spondylitis, chronic rheumatic heart disease, Crohn’s disease, haemorrhagic proctitis, multiple sclerosis, polymyalgia rheumatica, psoriasis, Reiters disease, rheumatic fever, sarcoidosis, ulcerative colitis, Wegener’s granulomatosis). We grouped autoimmune diseases into diseases with detectable and not detectable antibodies to distinguish between primarily antibody-driven and T-cell-driven autoimmune processes. Diabetes mellitus (type I or type II unspecified) was also examined. Sites of infectious/inflammatory diseases included respiratory (chronic sinusitis, influenza, nasopharyngitis/laryngitis, chronic bronchitis, pneumonia, tuberculosis), digestive (atrophic gastritis, cholangitis and cholecystitis, gastroenteritis, hepatitis), and genitourinary systems (chronic cystitis, chlamydia, venereal gonorrhoea, glomerulonephritis, pyelonephritis, nephrotic syndrome, syphilis, urethritis, orchitis and epididymitis, chronic prostatitis, genital herpes), systemic disease (meningitis, septicaemia, malaria, infectious mononucleosis), and others (herpes simplex, herpes zoster, mycoses, rickettsioses, poliomyelitis).

Statistical analysis. We compared the rates of brain cancer in men with and without immune-related conditions, modelled as a time-dependent variable (Preston, 2005), after adjusting for age (18–39, 40–49, 50–59, 60–69, 70–79, 80 + years), calendar year (1969–1974, 1975–1979, 1980–1984, 1985–1989 and 1990–1996), race (black, white), and number of hospital visits (1–2, 3–4, ≥5). Rate ratios (RRs) and 95\% confidence intervals (95\% CI) were calculated using Poisson regression (Breslow and Day, 1987; Preston, 2005).

To minimise the influence of surveillance bias (i.e., earlier detection of brain cancers among persons under close surveillance for other medical conditions) or reverse causation (i.e., undetected brain cancer causing or suppressing the manifestation of an immune-related condition), analyses were stratified by latency, defined as time between diagnosis of immune condition and end of follow-up (< 2, ≥2 years). The < 2/≥2 years choice was motivated by median survival time of patients with common glioma subtypes reported in the literature (CBTRUS, 2012). Differences between incidence rates for these two latency periods were evaluated using likelihood ratio tests. For major groups and specific conditions with sufficient numbers of cases, brain cancer risk was evaluated with number of conditions of interest (1, ≥2), finer categories of latency based on tertiles of cases (2–4, 5–9, ≥10 years), and stratified on age (< 50, 50–59, 60–69, ≥70 years) and race (black, white). Trend tests for latency and age were performed using ordinal variables defined on the basis of respective categorical variables.

P-values were two-sided and considered statistically significant if less than 0.05. For our secondary hypotheses (brain cancer associations with specific immune-related conditions), P-values were additionally adjusted for multiple comparisons testing using the false discovery rate ($P_{FDR}$) (Benjamini and Hochberg, 1995). Calculations were performed using the AMFIT module of Epicure (Version 1.4; HiroSoft International, Seattle, WA, USA).

RESULTS

We identified 4,383 cases of brain cancer (514 black patients and 3,869 white patients) among 4,501,578 patients with a mean...
follow-up time of 11.7 years (Table 1). Incidence rates of brain cancer were 5.2 and 9.0 per 100 000 person-years for blacks and whites, respectively. Median ages at entry into the study cohort for brain cancer cases were 53.1 years for blacks and 54.1 years for whites. Patients with brain cancer tended to be slightly older at study entry, had more hospital visits, and were followed for a shorter period of time than patients without brain cancer.

Risk of brain cancer in relation to prior immune-related conditions by latency (<2, ≥2 years) is shown for conditions for which the number of cases was ≥5 (Table 2). For all major groups including allergy/atopy, autoimmune diseases, diabetes, and infectious/inflammatory diseases, and for the majority of specific conditions, the RR for brain cancer was significantly higher for latency <2 years compared with latency ≥2 years, suggesting that some bias might contribute to the observed increase in risk. Therefore, in all future analyses, individuals whose immune-related conditions were diagnosed less than 2 years before brain cancer or end of follow-up had the corresponding person-years excluded. Following exclusion of the last 2 years of follow-up, allergic/atopic and autoimmune diseases (primarily with detectable antibodies) were generally associated with a non-significantly reduced risk of brain cancer, with the exception of Addison’s disease, which was associated with a significantly increased risk of brain cancer (RR = 2.47, 95% CI: 1.03, 5.94). After controlling for multiple comparisons testing, none of the associations with specific allergic/atopic or autoimmune conditions remained significant.

Infectious/inflammatory conditions as a group, diagnosed ≥2 years before brain cancer, was not associated with risk of brain cancer (RR = 1.00, 95% CI: 0.91, 1.10). However, atrophic gastritis (RR = 2.98, 95% CI: 1.65, 5.39), chronic prostatitis (RR = 1.67, 95% CI: 1.02, 2.74), and mycoses (RR = 1.29, 95% CI: 1.07, 1.54) were associated with an increased risk of brain cancer, whereas chronic bronchitis was associated with a reduced risk (RR = 0.78, 95% CI: 0.65, 0.93). After controlling for multiple comparisons testing, associations with atrophic gastritis and chronic bronchitis, but not other diseases, remained suggestive at borderline significance levels (P_{FDR} = 0.08 for both).

Rate ratios for brain cancer were further stratified by latency for major groups of immune-related conditions (Table 3). We found a significant trend of decreasing risk of brain cancer with longer latency of allergy/atopy (P = 0.02). Risk of brain cancer was also significantly reduced for patients with latency of allergy/atopy ≥10 years (RR = 0.60, 95% CI: 0.43, 0.83). Although the trend for risk of brain cancer with latency of diabetes did not reach significance (P = 0.15), history of diabetes with ≥10-year latency was associated with a significantly reduced risk of brain cancer (RR = 0.75, 95% CI: 0.62, 0.93). Autoimmune diseases without detectable autoantibodies, taken as a group, was associated with a significantly increased risk of brain cancer for latency 2–4 years (RR = 1.44, 95% CI: 1.02, 2.01), but not for latency ≥10 years (RR = 0.82, 95% CI: 0.55, 1.22), and the overall trend for latency was not significant.

The RRs are shown for number of immune-related conditions (Table 4). Risks of brain cancer tended to be more inversely associated with having ≥2 conditions than 1 condition, but the differences were not significant.

For allergy/atopy, autoimmune diseases, and infectious/inflammatory diseases with at least 50 cases, we saw few significant differences in RRs between patients <60 and ≥60 years of age or between blacks and whites (Table 5). An exception was chronic bronchitis, which was associated with a reduced risk of brain cancer in patients ≥60 years of age (RR = 0.69, 95% CI: 0.56, 0.87), but not in patients <60 years (RR = 0.98, 95% CI: 0.74, 1.31; P_{FDR} for interaction = 0.08).

**DISCUSSION**

The search for clues into the aetiology of brain cancer has uncovered few definitive risk factors. Previous studies of allergy/atopy and risk of glioma have found an inverse relationship (Calboli et al, 2011; Chen et al, 2011; Schlehofer et al, 2011), whereas cohort studies have often had limited power and shown inconsistent results (Schwartzbaum et al, 2003; Hagstromer et al, 2005; Turner et al, 2005; Hwang et al, 2012). Autoimmune conditions, diabetes, and infectious/inflammatory conditions have not been extensively examined in relation to brain cancer risk. Using a large cohort of black and white male US veterans with a relatively homogeneous socio-economic background, we were able to examine medically diagnosed major groups of immune-related conditions and specific immune-related conditions while minimizing potential biases associated with the case–control study design. We found allergy/atopy and diabetes of long latency to be associated with reduced risk of subsequent brain cancer. None of the associations with specific diseases withstand correction for multiple comparisons.

History of atopy/allergy has been associated with both increased and decreased risk of various cancers (Hwang et al, 2012; Rittmeyer and Lorenz, 2012). A review of 32 epidemiological studies on the subject found decreased risk for cancers of the brain, pancreas, and colon, but increased risk for lung and skin cancers. Inflammatory reactions accompanying atopy/allergy have been hypothesised to support carcinogenesis of some cancers, particularly of epithelial origin, whereas overall increased immune surveillance has been suggested as one potential explanation for decreased risk of other cancers (Turner et al, 2006). Some support

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**Table 1. Characteristics of brain cancer study cohort of white and black US male veterans**

| Characteristic                          | White | Non-case | Total | White | Non-case | Total | White | Non-case | Total |
|-----------------------------------------|-------|----------|-------|-------|----------|-------|-------|----------|-------|
| No. of patients                         | 3869  | 3 665.375| 3 669.244| 514   | 831 820  | 832.334| 4383  | 4 497 195| 4 501.578|
| No. of person-years                     | 27775 | 42 726.418| 42 754.193| 4011  | 9 883.084 | 9 887.095| 31 787 | 52 609.502| 52 641.289|
| Incidence rate (per 100 000 person-years)| 9     | 5.2      | 8.3    |       |          |       |       |          |       |
| Median age at entry, years              | 54.1  | 53.5     | 53.5   | 53.1  | 47.7     | 47.7   | 54    | 52.5     | 52.5   |
| Mean follow-up, years                   | 7.2   | 11.7     | 11.7   | 7.8   | 11.9     | 11.9   | 7.3   | 11.7     | 11.7   |
| Median age at brain cancer diagnosis, years | 61.8  | 61.2     | 61.7   |       |          |       |       |          |       |
| Median no. of hospital visits           | 4     | 3        | 3      | 4     | 3        | 3      | 4     | 3        | 3      |

Abbreviation: No. = number.
Table 2. Risk of brain cancer for immune-related conditions by latency (duration of follow-up) among US male veterans

| Conditiona | No. of BCs | RRb | 95% CI | No. BC | RRb | 95% CIc | LRT Pd |
|------------|-----------|-----|--------|--------|-----|---------|--------|
| **Total allergy/atopy** | | | | | | | |
| Allergic rhinitis | 51 | 1.51 | 1.15 | 1.99 | 141 | 0.85 | 0.72 | 1.01 <0.01 |
| Asthma | 19 | 1.51 | 0.96 | 2.37 | 46 | 0.8 | 0.6 | 1.07 0.03 |
| Excoria/dermatitis | 26 | 1.37 | 1.2 | 2.61 | 60 | 0.82 | 0.64 | 1.06 0.01 |
| Enzyma | 5 | 1.05 | 0.43 | 2.51 | 22 | 1.03 | 0.68 | 1.57 >0.50 |
| Urticaria | 3 | 1.84 | 0.59 | 5.7 | 6 | 0.65 | 0.29 | 1.44 0.17 |
| **Total autoimmune disease** | 60 | 1.45 | 1.13 | 1.88 | 215 | 0.96 | 0.84 | 1.1 0.01 |
| Autoantibodies detectable | 41 | 1.4 | 1.03 | 1.9 | 138 | 0.87 | 0.73 | 1.03 0.01 |
| Systemic involvement | 17 | 1.68 | 1.04 | 2.7 | 51 | 0.85 | 0.65 | 1.12 0.02 |
| Rheumatoid arthritis | 14 | 1.53 | 0.91 | 2.59 | 47 | 0.85 | 0.64 | 1.03 0.07 |
| Organ involvement | 29 | 1.41 | 0.98 | 2.03 | 90 | 0.86 | 0.7 | 1.07 0.03 |
| Addison's disease | 3 | 4.72 | 1.52 | 14.63 | 5 | 2.47 | 1.03 | 5.94 0.39 |
| Celiac disease | 2 | 1.05 | 0.43 | 2.51 | 22 | 1.03 | 0.68 | 1.57 0.50 |
| **Total infectious/inflammatory** | 180 | 1.49 | 1.28 | 1.73 | 559 | 1 | 0.91 | 1.1 >0.01 |
| Respiratory system | | | | | | | |
| Chronic sinusitis | 16 | 1.44 | 0.88 | 2.35 | 60 | 1.08 | 0.84 | 1.4 0.33 |
| Influenza | 9 | 2.3 | 1.2 | 4.42 | 10 | 0.99 | 0.81 | 1.17 0.01 |
| Nasopharyngitis/laryngitis | 0 | | | | | | |
| Chronic bronchitis | 53 | 3.08 | 1.2 | 7.38 | 129 | 0.78 | 0.65 | 0.93 <0.01 |
| Pneumonia | 84 | 1.48 | 1.19 | 1.83 | 186 | 0.99 | 0.86 | 1.16 <0.01 |
| Tuberculosis | 12 | 1.04 | 0.59 | 1.84 | 65 | 0.93 | 0.73 | 1.19 >0.01 |
| Digestive system | | | | | | | |
| Intestinal Infectious | 21 | 1.95 | 1.27 | 3 | 52 | 0.98 | 0.75 | 1.29 0.01 |
| Atrophic gastritis | 2 | 1.09 | 0.27 | 4.37 | 11 | 2.98 | 1.65 | 5.39 0.15 |
| Cholangitis/cholecystitis | 13 | 2.19 | 1.27 | 3.78 | 25 | 0.99 | 0.67 | 1.47 0.03 |
| Hepatitis | 3 | 0.81 | 0.26 | 2.53 | 15 | 0.82 | 0.49 | 1.37 >0.50 |
| Genitourinary system | | | | | | | |
| Glomerulonephritis | 4 | 0.61 | 0.23 | 1.62 | 11 | 0.7 | 0.39 | 1.26 >0.50 |
| Chronic pyelonephritis | 1 | 0.58 | 0.08 | 4.1 | 8 | 0.88 | 0.44 | 1.75 >0.50 |
| Nephrotic syndrome | 8 | 4.23 | 2.12 | 8.48 | 7 | 1.48 | 0.71 | 3.11 0.04 |
| Syphilis | 12 | 1.91 | 1.08 | 3.37 | 40 | 1.2 | 0 | 1.58 0.2 |
| Urethritis | 2 | 1.11 | 0.28 | 4.43 | 16 | 1.4 | 0.86 | 2.29 >0.50 |
| Orchitis/epididymitis | 13 | 2.39 | 1.06 | 5.45 | 53 | 1.29 | 0.98 | 1.69 0.43 |
| Chronic prostatitis | 6 | 1.69 | 0.76 | 3.77 | 16 | 1.67 | 1.02 | 2.74 >0.50 |
| Systemic | | | | | | | |
| Meningitis | 10 | 5.69 | 3.06 | 10.58 | 11 | 1.23 | 0.68 | 2.22 <0.01 |
| Septicaemia | 16 | 1.46 | 0.89 | 2.38 | 27 | 1.1 | 0.75 | 1.6 0.38 |
| Other | | | | | | | |
| Herpes simplex | 2 | 1.2 | 0.3 | 4.81 | 5 | 0.74 | 0.31 | 1.77 >0.50 |
| Herpes zoster | 3 | 0.9 | 0.29 | 2.78 | 8 | 0.8 | 0.4 | 1.6 >0.50 |
| Mycoses | 36 | 1.61 | 1.16 | 2.23 | 123 | 1.29 | 1.07 | 1.54 0.25 |
| Rickettsioses | 3 | 1.04 | 0.33 | 3.22 | 15 | 1.26 | 0.76 | 2.09 >0.50 |

Abbreviations: BC = brain cancer; CI = confidence interval; RR = rate ratio.

a Specific conditions shown for >5 cases.
b Adjusted for age category, calendar time, race, and number of hospital visits.
c After adjustment for multiple comparisons using the false discovery rate, no RRs for conditions with latency >2 years were significantly different from the null. False discovery rate P-values for chronic bronchitis (P = 0.08) and atrophic gastritis (P = 0.08) were suggestive.
d LRT P-value for difference in RR by latency <2 years and >2 years.
Immune conditions and brain cancer in US veterans

for immune surveillance theory comes from experimental animal studies, which found that circulating IgE may impede early tumour development (Jensen-Jarolim et al., 2008). However, at present, this theory cannot provide a uniform explanation for the range of associations observed for different cancers. Glioma is among the few cancers that exhibit more consistent inverse association with history of allergy/atopy. This is particularly intriguing because until recently the brain’s accessibility to various elements of immune system was considered limited due to the existence of the blood brain barrier (Mrass and Weninger, 2006). This view is changing, Our findings of reduced risk of brain cancer associated with prior allergic/atopic conditions are generally more modest than the results of previous studies and point to a more pronounced association with allergy of long rather than short duration (Wigertz et al., 2007; Turner et al., 2013). A recent meta-analysis evaluating the relationship between any allergy and risk of glioma found a combined OR of 0.60 (95% CI: 0.52, 0.69) (Chen et al., 2011). However, the cohort studies in this analysis contributed a small fraction of cases to the results. Insofar as case–control studies present the possibility of finding non-causal associations and trends through selection or recall bias, our findings provide further support of the association between allergy/atopy, particularly of long latency, and risk of brain cancer. The weaker RRs with allergy/atopy found in our study compared with other studies, which combined men and women, may be potentially explained by effect modification by sex. Two nested case–control studies using pre-diagnostic levels of serum IgE found greater reductions in risk of glioma among women (Schlehofer et al., 2011; Schwartzbaum et al.,

Table 3. Risk of brain cancer for selected conditions by latency (duration of follow-up) among US male veterans

| Condition* | 2–4 years | 5–9 years | ≥10 years |
|------------|-----------|-----------|-----------|
|            | No. of BCs | RRb       | 95% CI    | No. of BCs | RRb       | 95% CI    | No. of BCs | RRb       | 95% CI    | LRT P<sup>c</sup> (trend) |
| Total allergy/atopy | 48 | 0.96 | 0.7 | 1.3 | 58 | 1.01 | 0.8 | 1.3 | 35 | 0.6 | 0.4 | 0.8 | 0.02 |
| Total autoimmune disease | 71 | 1.07 | 0.9 | 1.4 | 74 | 0.96 | 0.8 | 1.2 | 70 | 0.87 | 0.7 | 1.1 | 0.38 |
| Autoantibodies detectable | 39 | 0.83 | 0.6 | 1.1 | 48 | 0.88 | 0.7 | 1.2 | 51 | 0.89 | 0.7 | 1.2 | 0.13 |
| Systemic involvement | 17 | 0.98 | 0.6 | 1.6 | 17 | 0.82 | 0.5 | 1.3 | 17 | 0.78 | 0.5 | 1.3 | 0.2 |
| Organ involvement | 23 | 0.73 | 0.5 | 1.1 | 32 | 0.9 | 0.6 | 1.3 | 35 | 0.94 | 0.7 | 1.3 | 0.26 |
| Autoantibodies not detectable | 34 | 1.44 | 1.2 | 1.6 | 31 | 1.12 | 0.8 | 1.6 | 24 | 0.82 | 0.6 | 1.2 | >0.50 |
| Diabetes | 143 | 1.02 | 0.9 | 1.2 | 159 | 1.06 | 0.9 | 1.3 | 96 | 0.75 | 0.6 | 0.9 | 0.15 |
| Total infectious/inflammatory disorders | 189 | 1.12 | 1.1 | 1.3 | 186 | 0.99 | 0.9 | 1.2 | 184 | 0.91 | 0.8 | 1.1 | >0.50 |

Abbreviations: BC = brain cancer; CI = confidence interval; LRT = likelihood ratio test; RR = rate ratio.
*Excludes cases with brain cancer in first 2 years.
*bAdjusted for age, calendar time, race, and number of hospital visits.
*cLRT P-value used treated categorical latency variables as ordinal.

Table 4. Risk of brain cancer for number of conditions within a condition category among US male veterans

| Condition* | 1 condition | ≥2 conditions |
|------------|-------------|--------------|
|            | No. of BCs | RRb       | 95% CI    | No. of BCs | RRb       | 95% CI    | LRT P<sup>d</sup> |
| Total allergy/atopy | 131 | 0.86 | 0.72 | 1.02 | 10 | 0.77 | 0.42 | 1.44 | 0.14 |
| Total autoimmune disease | 192 | 1 | 0.86 | 1.16 | 23 | 0.72 | 0.48 | 1.09 | 0.25 |
| Total infectious/inflammatory disorders<sup>a</sup> | 445 | 1.1 | 0.95 | 1.16 | 114 | 0.85 | 0.7 | 1.02 | 0.11 |

Abbreviations: BC = brain cancer; CI = confidence interval; LRT = likelihood ratio test; RR = rate ratio.
*aExcludes cases with brain cancer in first 2 years.
*bAdjusted for age, calendar time, race, and number of hospital visits.
*cLikelihood ratio test P-value for difference in RR by number of conditions, categories 1 and ≥2 conditions.
*dDoes not include diabetes.

Excludes cases with brain cancer in first 2 years.
IgE and risk of glioma (Calboli 2012). However, sex differences were not found in another study of type II diabetes (T2DM) involves dysregulation in various immunological pathways (elevated levels of interleukin (IL)-1β, IL-6, the NF-κB, and JNK) contributing to its pathogenesis (Donath and Shoelson, 2011). Finally, non-diifferential exposure misclassification, particularly for common conditions such as allergic rhinitis that do not generally have serious clinical implications or be negligible.

Chronic inflammation and infections have been associated with increased risk for multiple cancers, possibly due to a deficient immune response (Hopton Cann et al, 2006; Scheurer et al, 2008; Conti et al, 2010). Although we found significant associations for brain cancer with several inflammatory conditions, none withstood correction for multiple comparisons. Replication of these specific associations in future studies is warranted.

Several strengths and limitations should be considered in the interpretation of our results. Our study had 4383 cases of brain cancer among men, including 514 black men. The follow-up time of our study was reasonably long (average, 11.7 years). These large numbers allowed us to explore variation in risk of brain cancer for more common immune-related conditions by latency, age, and race. The potential for confounding by socio-economic status was limited, because patients within the VA system typically have lower socio-economic status (Randall et al, 1987), and previous VA studies found similar health care utilisation rates and outcomes for blacks and whites (Deswal et al, 2004).

A major strength of our study was that immune-related conditions were ascertained through medical diagnoses from hospital discharge records. Because the diagnoses of interest were

### Table 5. Risk of brain cancer for selected conditions by age and race among US male veterans

| Condition                      | <60 years | ≥60 years | Whites | Blacks |
|-------------------------------|----------|----------|--------|--------|
|                               | RR       | 95% CI   | RR     | 95% CI | RR     | 95% CI | LRT P* |
| Total allergy/atopy           | 0.81 0.62 1.07 | 0.87 0.71 1.08 | >0.50 | 0.88 0.73 1.04 | 0.67 0.39 1.17 | 0.42 |
| Eczema/dermatitis             | 0.76 0.5 1.16 | 0.87 0.63 1.2 | >0.50 | 0.8 0.61 1.05 | 1.03 0.51 2.07 | 0.49 |
| Total autoimmune disease      | 0.95 0.76 1.19 | 0.96 0.81 1.14 | >0.50 | 0.97 0.84 1.12 | 0.87 0.55 1.37 | >0.50 |
| Autoantibodies detectable     | 0.92 0.7 1.21 | 0.84 0.67 1.04 | 0.27 | 0.87 0.73 1.04 | 0.87 0.51 1.48 | >0.50 |
| Systemic involvement          | 0.97 0.61 1.52 | 0.79 0.56 1.12 | 0.28 | 0.83 0.62 1.11 | 1.13 0.5 2.35 | 0.41 |
| Organ involvement             | 0.85 0.6 1.19 | 0.87 0.67 1.13 | >0.50 | 0.89 0.71 1.1 | 0.69 0.35 1.4 | >0.50 |
| Autoantibodies not detectable | 0.9 0.63 1.29 | 1.25 0.96 1.62 | 0.13 | 1.14 0.92 1.42 | 0.74 0.31 1.79 | 0.31 |
| Diabetes                      | 0.86 0.71 1.06 | 0.99 0.87 1.11 | >0.50 | 0.96 0.86 1.08 | 0.89 0.68 1.17 | >0.50 |
| Total infectious/inflammatory disorders | 0.95 0.82 1.09 | 1.04 0.93 1.17 | 0.35 | 0.99 0.9 1.09 | 1.08 0.86 1.36 | 0.34 |
| Chronic sinusitis             | 0.84 0.53 1.32 | 1.25 0.92 1.7 | 0.14 | 1 | 0.76 1.32 | 1.88 1.01 3.53 | 0.08 |
| Chronic bronchitis            | 0.98 0.74 1.31 | 0.69 0.56 0.87 | <0.01 | 0.79 0.66 0.95 | 0.67 0.35 1.3 | >0.50 |
| Pneumonia                     | 1.06 0.82 1.36 | 0.96 0.8 1.15 | 0.2 | 0.99 0.84 1.16 | 1.01 0.7 1.44 | >0.50 |
| Tuberculosis                  | 0.95 0.64 1.4 | 0.91 0.66 1.25 | >0.50 | 0.88 0.66 1.17 | 1.11 0.69 1.78 | 0.34 |
| Intestinal infectious         | 0.78 0.49 1.24 | 1.14 0.81 1.6 | 0.18 | 1.04 0.79 1.38 | 0.5 0.16 1.57 | 0.19 |
| Orchitis/epididymitis         | 0.91 0.54 1.55 | 1.52 1.11 2.08 | 0.11 | 1.18 0.87 1.61 | 1.86 1.05 3.3 | 0.15 |
| Mycoses                       | 1.28 0.98 1.66 | 1.29 1.01 1.66 | >0.50 | 1.32 1.09 1.6 | 1.1 0.67 1.82 | 0.46 |

Abbreviations: CI = confidence interval; LRT = likelihood ratio test; RR = rate ratio.
*a*Adjusted for continuous age, calendar time, race, and number of hospital visits.
*b*Adjusted for age, calendar time, and number of hospital visits.
*c*Excludes cases with brain cancer in first 2 years. Specific conditions shown for >50 cases.
*d*Likelihood ratio test; RR = rate ratio.
*e*Likelihood ratio test; P-value for difference in RR by age <60 years and ≥60 years. After adjustment for multiple comparisons for specific conditions using the false discovery rate, no LRT P-values were significant. The false discovery rate P-value for chronic bronchitis (P = 0.08) was suggestive.
*f*Likelihood ratio test; P-value for difference in RR by race. After adjustment for multiple comparisons for specific conditions using the false discovery rate, no LRT P-values were significant.
recorded by medical staff before, and irrespective of, the outcome, the reported associations are not affected by recall or related biases as may occur in case–control studies. However, as medical attention for an immune-related condition might result in earlier brain cancer detection, or the presence of asymptomatic brain cancer can change immune function (Waziri, 2010), surveillance bias and reverse causality remain as possibilities. Glia has been shown to illicit strong immunosuppressive effects (Wu et al, 2009; Di Tomaso et al, 2010) so that reverse causation bias, if present, would manifest in reduced risks for immune-related conditions. Instead, the increased RRs that we found for a wide range of immune-related conditions with short latency (< 2 years) suggest that surveillance bias is more likely to be present. In addition, if non-specific immunosuppression were having a major role, we would expect lower RRs with shorter latency; yet we observed the lowest RRs with longest duration of allergy and diabetes. Glioblastoma patients, who make up the largest group of brain cancer diagnoses in the United States, especially at older ages, have a median survival time of approximately 12–14 months, so restricting analyses to latency ≥ 2 years should have minimised the potential for surveillance bias (CBTRUS, 2012). Other glioma subtypes have a somewhat better prognosis (e.g., lower grade astrocytoma) so that this bias may still affect our results, particularly in younger age groups.

Limitations of this study include the lack of information on histological type of brain cancer, potential risk factors such as smoking, alcohol, diet, physical activity and body mass, radiation exposure, occupation, genetic predisposition, diagnostic testing, laboratory findings, and treatment data (e.g. anti-inflammatory or antihistamine medications). However, evidence concerning associations with intake of antihistamine or anti-inflammatory medication and risk of brain cancer is inconsistent. Use of antihistamine medication has been associated with both increased (Scheurer et al, 2008, 2011; Amirian et al, 2013) and decreased (Schlehofer et al, 1999; Schoemaker et al, 2006; McCarthy et al, 2011) risk of glioma, whereas use of NSAIDs has been associated with reduced risk of glioma in some studies (Sivak-Seals et al, 2004; Scheurer et al, 2008, 2011; Ferris et al, 2012), but not others (Daugherty et al, 2011; Bannon et al, 2013; Gaist et al, 2013). We also did not have information on diagnostic or therapeutic radiation procedures, which are the primary sources of ionising radiation exposure in the general US population. However, we excluded patients with prevalent cancer and patients who did not survive or were diagnosed with cancer within the first year of index admission, minimising the possible role of therapeutic radiation exposure. Moreover, radiation is not typically used for diagnosis or treatment of the immune-related conditions and, therefore, is unlikely to confound the associations of interest.

As our analyses are based on hospital discharge records, it is possible that there was some misclassification of exposure. A reliability study of VA data over a 3-month period in 1995 by Kashner (1998), reported that the VA’s Patient Treatment File (PTF) contained extra diagnoses per discharge when compared with medical charts for certain conditions, including diabetes. However, diabetes mellitus was among the conditions with the best agreement between the PTF and medical records with respect to secondary diagnoses. In our case, less severe conditions (e.g., allergic rhinitis) may have been under-reported, but we have no reason to believe this was differential by brain cancer diagnosis, particularly for conditions with long latency. Our patients had access to standardised medical care; however, persons who use the VA health care system are not restricted from using other health-care systems. For example, as the VA system does not have many emergency rooms, VA patients who require acute care might not report to a VA hospital. However, it is likely that persons with chronic conditions would continue their care at the VA, as it would be without cost. Similarly, if an individual reported to an emergency room due to their having cancer, they would likely be referred back to the VA system.

Persons who utilise the VA medical system have been reported to be of lower socio-economic status (as mentioned above) and in poorer health compared with the general population, suggesting that caution should be exercised when extrapolating these results to the overall US male population (Agha et al, 2000). However, it is encouraging that the rates of brain cancer in this cohort (5.2 per 100 000 person-years for blacks and 9.0 per 100 000 person-years for whites) are similar to the incidence rates in the Surveillance, Epidemiology, and End Results Program (1973–2009) (4.4 per 100 000 person-years for black males and 9.7 per 100 000 person-years for white males for years 1973–1996 in the same age range). This suggests that under-ascertainment of brain cancer diagnoses and treatment outside of the VA system was not extensive. The rarity of certain immune-related conditions, particularly in men, limited statistical power of some analyses and led to imprecision in risk estimates.

In summary, we found history of allergy/atopy and diabetes of long latency (10 years or more) to be associated with reduced risk of subsequent brain cancer. The findings in this large cohort with disease ascertainment based on hospital records, generally lend support to an inverse association between immune-related conditions and subsequent brain cancer risk demonstrated in case–control studies. To improve understanding of mechanisms underlying these associations, future studies with prospectively ascertainment and validated medical diagnoses, treatment information, and biological samples to measure a variety of immune function markers are needed.

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