The PAU survey: measurements of the 4000 Å spectral break with narrow-band photometry

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ABSTRACT
The D4000 spectral break index is one of the most important features in the visible spectrum, as it is a proxy for stellar ages and is also used in galaxy classification. However, its direct measurement has always been reserved to spectroscopy. Here, we present a general method to directly measure the D4000 with narrow-band (NB) photometry; it has been validated using realistic simulations, and then evaluated with PAUS NBs, cross-matched with VQuaS spectra \( (i_{AB} < 22.5, 0.562 < z < 0.967) \). We also reconstruct the D4000 with the SED-fitting code CIGALE; the use of PAUS NBs instead of broad-bands significantly improves the SED fitting results. For \( D_{4000} \), the direct measurement has \( \langle \text{SNR} \rangle \sim 4 \), but we find that for \( i_{AB} < 21 \) all direct D4000 measurements have \( \text{SNR} > 3 \). The CIGALE \( D_{4000} \) has \( \langle \text{SNR} \rangle \sim 20 \), but underestimates the error by >50 per cent. Furthermore, the direct method recreates the D4000–SFR relation, as well as the D4000–mass relation for blue galaxies (for red galaxies, selection effects impact the results). On the other hand, CIGALE accurately classifies galaxies into red and blue populations. We conclude that the direct measurement of D4000 with narrow-band photometry is a promising tool to determine average properties of galaxy samples, with results compatible with spectroscopy.

Key words: galaxies: evolution
filters of the order of FWHM $\sim$ 100 Å. These offer much larger spectral resolution and higher redshift precision than their broad-band counterparts, without increasing observation time as much as spectroscopy. They also keep other advantages of photometric surveys: no specific target selection is required, and the reduction of their images is still significantly less complex than that of spectra. Most existing narrow-band surveys so far observed small areas of the order of few deg$^2$, covering the visible spectrum with medium-width filters of FWHM $> 200$ Å, such as COMBO-17 (Wolf et al. 2003) or ALHAMBRA (Moles et al. 2008). Also, surveys specifically designed for the observation of certain emission lines with very few filters have also been carried out, like SILVERRUSH (Ouchi et al. 2018) for the Ly$\alpha$ line.

Narrow-band surveys covering significant angular areas, sampling the whole visible spectrum with filters of FWHM $\sim$ 100 Å, have been proposed during the past decade, and are currently starting to reveal its potential. One example of these surveys is the ongoing Physics of the Accelerating Universe Survey (PAUS; Benítez et al. 2009), whose data is employed in this work. PAUS has been primarily designed to determine high-precision photometric redshifts (Marti et al. 2014; Eriksen et al. 2019), mostly allowing for precise cross-correlation of lensing and redshift distortion probes (Gaztañaga et al. 2012). Other examples of narrow-band surveys with larger angular coverage are J-PAS (Benítez et al. 2014), J-PLUS (Cenarro et al. 2019), and S-PLUS (De Oliveira et al. 2019). These more general-purpose narrow-band surveys show potential to resolve specific spectral features, with closer approximation to spectroscopy (e.g. Stothert et al. 2018; Martínez-Solá et al. 2020).

The observation of specific spectral features besides the spectral continuum allows us to determine key galaxy properties such as star formation rate (SFR), stellar age, and metallicity (e.g. Worthey 1994; Kauffmann et al. 2003a); these spectral features are the results of different physical processes, and thus are a proxy to study the formation and evolution of galaxies. Generally, emission lines provide information about the ionized interstellar medium (ISM; Kewley, Nicholls & Sutherland 2019), while absorption lines inform of the properties of the stellar population (Maraston et al. 2009). For example, emission lines are commonly used to select active galactic nuclei (AGNs) based on the Baldwin–Phillips–Terlevich (BPT) diagram (Baldwin, Phillips & Terlevich 1981; Kauffmann et al. 2003c; Kewley et al. 2013), since they allow us to discern between ionization in the intergalactic medium (IGM) due to a soft UV source (e.g. star formation) or hard UV (e.g. AGN emission). On the other hand, absorption Lick indices are proxies for inferring stellar ages and metallicities (Worthey 1994; Thomas, Maraston & Bender 2003).

The direct measurement of spectral features is generally reserved to spectroscopic surveys. However, galaxy properties can also be estimated with other methods, where the entirety of the spectrum is fit to stellar population models (full spectrum fitting). These fits are often convolved with the line-of-sight velocity distribution, which makes it especially useful to study galaxy kinematics. Some examples of full spectrum fitting codes are STECKMAP (Ocvirk et al. 2006), ULYSS (Koleva et al. 2009), and PPEX (Cappellari 2017). Perhaps the most prominent one is spectral energy distribution (SED) fitting, in which a linear combination of spectral templates associated with different physical systems (stellar population models, dust emission/attenuation, AGNs, etc.) are fit to the observed SED, and the galaxy properties are inferred from the fit. This is a method widely applied to broad-band photometry, and several codes exist in the literature for this purpose, such as LEPHARE (Ilbert et al. 2006), CIGALE (Noll et al. 2009; Boquien et al. 2019), or PROSPECT (Robotham et al. 2020); moreover, the application of SED-fitting codes for narrow-band photometric surveys has already been evaluated in Delgado et al. (2021) for the mini J-PAS data release.

One spectral feature of special interest for this work are spectral breaks: combinations of absorption lines in broader features that damp the continuum emission at given wavelength ranges. The most important of these features in the optical part of galaxy spectra is the 4000 Å break, generated around that wavelength range by the absorption of several ionized metallic elements, as well as the latest lines in the Balmer series. The strength of this spectral break is quantified by the ratio between the continuum fluxes before and after the break itself; this quantity is known as the D4000 spectral index, or just D4000. Two definitions for the D4000 with slightly different wavelength ranges exist in the literature: the original D4000 wide definition (Brualdi 1983; Hamilton 1985), and another defined in a narrower wavelength span (Balogh et al. 1999), which is the current standard. Henceforth, we will refer to these definitions as D4000, and D4000, respectively.

While D4000 is generally preferred, mostly because it is far less sensitive to reddening effects due to the smaller continuum regions, both D4000 definitions have been extensively applied to study galaxy properties (e.g. stellar age, metallicity). The D4000 strength is correlated to the age of stellar populations: as galaxies grow old, they become redder and the strength of D4000 increases. However, it is also correlated with metallicity (Worthey 1994), and when D4000 is used to determine galaxy ages, it is generally combined with metallicity indicators to break the age–metallicity degeneracy.

Given these characteristics, the D4000 has been widely used in galaxy evolution works as a measurement of stellar age, based mostly on D4000 (e.g. Balogh et al. 1999; Brinchmann et al. 2004; Marcillac et al. 2006; Siudek et al. 2017), but also D4000 (e.g. Bruzual 1983; Mignoli et al. 2005; Kriek et al. 2011; Kim et al. 2018), or even alternative definitions (Tresse et al. 1999). In addition to these studies of galaxy properties, the D4000 has also been used in cosmology to independently determine the Hubble expansion rate, $H(z)$, with the method of cosmic chronometers (Moresco et al. 2012).

In that work, the measurement of D4000 is used to infer stellar ages of passive red galaxies; the age difference between redshift bins allows us to directly estimate the evolution of the Hubble parameter.

While the D4000 has been traditionally measured with spectroscopic observations, narrow-band photometric surveys may also yield enough spectral resolution for its measurement (Stothert et al. 2018; Anthopho, Ferreras & Silk 2020). In this work, we will assess the possibility of measuring the D4000 with PAUS narrow-band photometry. In fact, due to its high spectral resolution for an imaging survey, several works have already partially explored the possibilities of its photometry to detect spectral features and/or derive galaxy properties with PAUS. For example, in Johnston et al. (2021) and Tortorelli et al. (2021) CIGALE is applied to PAUS photometry to derive rest-frame colours and luminosities, as well as stellar masses and SFRs, while the photometric redshift code developed in Alarcon et al. (2021) also estimates emission line fluxes. Detection of spectral features has also been evaluated, like the Ly$\alpha$ line for intensity mapping studies in Renard et al. (2021), or even the possibility of D4000 measurements briefly evaluated with mock catalogues and synthetic data sets in Stothert et al. (2018), which is a precursor to the work presented here.

In order to evaluate PAUS capabilities to determine the D4000 value, we develop an estimator for directly measuring this spectral feature over its narrow-band photometry, and compare this methodology against the spectroscopic measurements provided by the VIMOS Public Extragalactic Redshift Survey (VIPERS), as well
as the reconstruction of the D4000 values obtained from the SED-fitting code CIGALE. For clarity, we will refer throughout this work to the D4000 value measured with PAUS narrow-band photometry as photometric measurement, the D4000 measurement over VIPERS spectra as spectroscopic measurement, and the value inferred from CIGALE SED fitting as reconstruction.

These estimation methods will be compared by evaluating the D4000 distribution and its average trends versus different observables (redshift, stellar mass, SFR). The validity of these D4000 estimations to discriminate between red and blue galaxies when compared to a fiducial classification will also be assessed. It is worth noting that this is the first work where the D4000 is determined by CIGALE using narrow-band photometry and compared to spectroscopy, although other works using CIGALE-derived D4000 with only broad-band photometry already exist. For example, in Buat et al. (2011) it is studied along other spectral features, in Boquien et al. (2012) it is used to examine the dependence of physical parameters on attenuation, and in Johnston et al. (2015) it is employed for galaxy classification. Moreover, the reconstruction of the D4000 (as well as other spectral indices) with SED-fitting techniques has also been evaluated in Mejía-Narváez et al. (2017) with synthetic J-PAS data.

The paper is structured as follows. In Section 2, we will explain the observational data sets that have been used. The D4000 estimator specifically developed for narrow-band photometry will be discussed in Section 3, as well as the reconstruction of D4000 and other galaxy properties with CIGALE. In Section 4, the performance of the D4000 estimator will be evaluated against spectroscopic and SED-fitting results by evaluating its SNR, distribution and bias. In Section 5, the evolution of the D4000 as determined by each estimation method will be examined, as well as its performance for galaxy classification. We will conclude with Section 6.

2 DATA SAMPLES

2.1 PAUS

PAUS is a narrow-band photometric survey carried out at the William Herschel Telescope. Observations are performed with a special-purpose instrument, the PAU Camera (Padilla et al. 2019), with a configuration of 40 narrow-band filters (NBs) uniformly spaced between 4550 Å and 8450 Å in steps of 100 Å (see Fig. 1). This filter configuration yields an average spectral resolution of $R \sim 65$. It has targeted five different fields: the COSMOS field (Laigle et al. 2016) and the wide fields (W1, W2, W3, W4) of the Canada–France–Hawaii Telescope Legacy Survey (CFHTLenS; Cuillandre et al. 2012; Hudelot et al. 2012). Photometry in all these fields is performed based on a reference broad-band catalogue (COSMOS and CFHTLS, respectively). For each object, an aperture mask is computed using the coordinates of the reference catalogue, and convolving the shape of the object (point source for stars, Sérsic profile for galaxies) with the point spread function (PSF) of the single-epoch image. Fluxes are integrated in these aperture masks up to a certain percentage of the estimated total flux of the object (a process known as forced photometry, see Eriksen et al. 2019, Cabayol et al. 2021 and Serrano et al. 2022).

The current angular footprint of the survey is ~50 deg$^2$, and its catalogue is complete up to $i_{AB} \leq 23$, with slightly over 3 million galaxies if all fields are considered. For all these galaxies, the computation of photo-z is currently restricted for objects in the range $0 < z < 1.2$, hence all objects at $z > 1.2$ are placed at lower redshifts in current photo-z runs (work is in progress to extend redshift calculations up to $z < 2$). These $z > 1.2$ objects accounts for 5 percent of the objects approximately, however, it does not affect the results in our paper, as the redshift cuts in our selected sample exclude any objects with spec-z above 1.2 (see Section 2.3).

Early results show that the photometric redshift is determined with an error of $\sigma_{\text{ph}}/(1+z) = 0.0037$ to $i_{AB} < 22.5$ for a 50 percent redshift quality cut (Eriksen et al. 2019) in the COSMOS field. These benchmark results have been improved with enhancements over the original photo-z code (Alarcon et al. 2021), or machine learning approaches (Eriksen et al. 2020; Soo et al. 2021). In all these works (as well as in this paper), photo-z is computed using both PAUS NBs and the broad-band photometry of the reference catalogues (CFHTLS for the W1 field).

2.2 VIPERS

The VIMOS public extragalactic redshift survey (VIPERS; Guzzo et al. 2013; Scodeggio et al. 2018) is a spectroscopic survey carried out at the Very Large Telescope in Cerro Paranal, Chile, with the VIMOS instrument (LeFevre et al. 2003). The grism used for this survey (low-resolution red) yields a spectral resolution $R \approx 220$ in a wavelength range of 5500–9500 Å (light-grey shaded area in Fig. 1). For this work, we utilize its second public data release, PDR-2 (Scodeggio et al. 2018).

The target selection of PDR-2 comprises objects in the redshift interval $0.5 < z < 1.2$, with a magnitude limit of $i_{AB} \leq 22.5$, which makes it one of the deepest spectroscopic catalogues publicly available. It is worth noting that deeper spectroscopic surveys do exist in fields overlapping with PAUS, such as VANDELS (Pentericci et al. 2018), VUDS (Le Fèvre et al. 2015), VVDS (Le Fèvre et al. 2013), or zCOSMOS-deep (Lilly et al. 2006), just to mention the surveys employing the same spectrograph as VIPERS. However, VIPERS has the largest angular coverage of them, as well as the closest magnitude limit to PAUS ($i_{AB} < 22.5$ versus $i_{AB} < 23$). Moreover, at its redshift range ($0.5 < z < 1.2$), VIPERS also offers an unique combination of sampled volume and galaxy density (almost 90 000 galaxies, see Scodeggio et al. 2018). Hence, it can be considered the closest catalogue to a perfect PAUS spectroscopic counterpart. The complete catalogue covers a total of ~23.5 deg$^2$, split between the CFHTLS W1 and W4 fields, which results in a catalogue of 86 775 galaxies in total.
Figure 2. Survey footprint in the W1 field for PAUS (orange), VIPERS (purple), and the cross-match between both (black). Each point represents an object in the catalogue.

Table 1. Wavelength ranges for the blue and red rest-frame bands of equation (1), for both the D4000\textsubscript{w} and D4000\textsubscript{r} definitions.

|          | Blue band (Å) | Red band (Å) |
|----------|---------------|---------------|
| D4000\textsubscript{w} | 3750–3950     | 4050–4250     |
| D4000\textsubscript{r} | 3850–3950     | 4000–4100     |

2.3 Sample selection

Thanks to its unprecedented volume and the wealth of auxiliary data, VIPERS is an ideal spectroscopic data set to validate PAUS capabilities. We have cross-matched the W1 VIPERS catalogue with PAUS observations, using a radius of 1 arcsec to match objects between catalogues. We have not considered the W4 field, as the current coverage of PAUS in W4 is significantly smaller than W1 (approximately 2 deg\textsuperscript{2} versus 14 deg\textsuperscript{2}), and very uneven between narrow bands. For these reasons, no photo-z calculation in W4 has been carried out yet at the time of writing.

In Fig. 2, the footprints of both catalogues in the W1 field are displayed; the cross-match between both yields a sample of 33 363 objects covering the whole PAUS redshift range (0 < z < 1.2). Moreover, we applied a cut based on the VIPERS redshift quality flags: only objects with 2 ≤ zflag ≤ 10 or 22 ≤ zflag ≤ 30 were included in the sample (Scoddogio et al. 2018). This cut keeps all the primary (i.e. in the original target catalogue) and secondary (i.e. observed in the slit but not in the target catalogue) objects with a redshift measurement good enough for science (> 90 per cent confidence level), while removing AGNs from the sample. This redshift quality cut leaves the selected sample with 28 788 galaxies.

Given that VIPERS spectra are observed in a redder wavelength span than PAUS, the wavelength range selected for D4000 measurements goes from 5850 Å to 8350 Å (Fig. 1, dark-grey area). With this choice of wavelength coverage, the low SNR regions at the limits of the spectrograph range are avoided, and an extra PAUS NB is left outside of range in the red end to avoid undesired boundary effects in the D4000 estimation. This restriction in wavelength coverage also implies a constraint in the redshift interval where the D4000 can be measured, as shown in Fig. 1 by the horizontal black lines. For D4000\textsubscript{w} and D4000\textsubscript{r}, with wavelength ranges defined in Table 1, this redshift range is 0.562 < z < 0.967 and 0.521 < z < 1.039, respectively. Nevertheless, we have applied just the most stringent cut of 0.562 < z < 0.967, in order to use the same sample to evaluate both D4000\textsubscript{w} and D4000\textsubscript{r}. After applying this redshift cut to both PAUS photo-z and VIPERS spec-z, the selected sample has 17 375 objects, with mean redshifts ⟨z\textsubscript{PAUS}⟩ = 0.731 and ⟨z\textsubscript{VIPERS}⟩ = 0.743 (see Fig. A1). This slight discrepancy in mean redshift may be due to the redshift focusing artefacts in the PAUS photo-z distribution, which appear as horizontal bands in Fig. A1. This effect is stronger for VIPERS objects at z > 0.7, where some galaxies are wrongly assigned a PAUS photo-z in the range 0.7 < z < 0.85, thus lowering the mean redshift of the whole distribution. Nevertheless, despite these artefacts the overall PAUS photo-z performance is significantly better than the CFHTLS broadband photo-z, as discussed in Section A.

In addition to the redshift cut, once the D4000 has been computed all objects with an invalid D4000 measurement have been removed. This includes objects with missing NBs or masked spectral regions needed for the photometric D4000 calculation, as well as objects whose CIGALE fits were deemed not good enough. This leaves a final sample of 17 241 galaxies that is used throughout this work.

3 METHODOLOGY: D4000 ESTIMATORS

In this section, we will describe in detail the two different methods used to measure the D4000 from PAUS data: the photometric measurement by determining the D4000 rest-frame bands using PAUS NBs (Section 3.1), and the reconstruction with the CIGALE SED-fitting that, among other physical parameters, yields the D4000 of the galaxies in the selected sample (Section 3.2). Moreover, in Section 3.1 we will also briefly review the general D4000 definition, as well as its calculation with spectroscopic data.

3.1 Direct estimator

3.1.1 Estimator definition

The measurement of the D4000 is defined as the ratio between two average flux frequency densities (F\textsubscript{v}), integrated in rest frame (also known as rest-frame bands),

\[
D4000 = \frac{\langle F_v \rangle_{\text{red}}}{\langle F_v \rangle_{\text{blue}}} . \tag{1}
\]

If these average flux density bands are to be directly measured on spectra, its integration is straightforward,

\[
\langle F_v \rangle = \frac{1}{\lambda_{\text{max}} - \lambda_{\text{min}}} \int_{\lambda_{\text{min}}}^{\lambda_{\text{max}}} (1 + z) d\lambda F_\nu(\lambda), \tag{2}
\]

where \(\lambda_{\text{min}}\) and \(\lambda_{\text{max}}\) are the wavelength boundaries for a given band, and \(z\) is the redshift of the object.
In equation (1), \( \langle F_r \rangle_{\text{red}} \) is the integrated flux of the rest-frame band defined at wavelengths larger than 4000 Å (i.e. the region unaffected by the combination of absorption lines that causes the break itself), and \( \langle F_r \rangle_{\text{blue}} \) is the integrated flux at the rest-frame band at shorter wavelengths (where the continuum is affected by the set of absorption lines of the 4000 Å break). The wavelength boundaries of these rest-frame bands are specified in Table 1, as specified in Bruzual (1983) for D4000\(_w\) and Balogh et al. (1999) for D4000\(_b\). We will refer to these bands, as well as to their integrated fluxes, as red band and blue band, respectively (or simply D4000 bands). The stronger the spectral break (i.e. the more absorption), the fainter the blue band compared to the red one, and the higher the D4000 value (since in equation (1) the flux of the red band is in the numerator, while the blue band flux is in the denominator).

In this work, the spectroscopic measurements of the D4000 from VIPERS spectra have been computed following equation (2) with the EZ redshift code (Garilli et al. 2010). If the D4000 has to be determined with narrow-band photometry, however, equation (2) cannot be directly applied. The spectral resolution of the NBs is of the same order of magnitude as the width of the D4000 bands, and thus common numerical integration methods (e.g. Trapezoidal rule, Simpson’s rule) may result in large inaccuracies. Consequently, we have developed a specific estimator to compute the D4000 bands with narrow-band photometry. While its performance will only be evaluated for PAUS in this work, by definition it can be applied to any combination of narrow-band filters, as long as their FWHM is smaller than the D4000 bands when shifted to the observed frame, and the wavelength coverage in the redshift range of interest is reasonably complete.

For PAUS NBs, \( \langle F_v \rangle \) is computed as

\[
\langle F_v \rangle = \frac{\sum \Delta \lambda \lambda_i^2 \sigma_i^{-2} F_{v_i}}{\sum \Delta \lambda \lambda_i^2 \sigma_i^{-2}},
\]

where \( F_{v_i} \) and \( \sigma_i \) are the flux and its error for the band \( i \), respectively, and the sum is over all the NBs considered for the photometric D4000 measurement (Fig. 1).

The \( r_i \) factor is the following ratio

\[
r_i = \frac{\int_{\lambda_{\text{min}}(1+z)}^{\lambda_{\text{max}}(1+z)} d\lambda \ R_i(\lambda)}{\int_0^{\lambda_{\text{max}}} d\lambda \ R_i(\lambda)},
\]

where \( R_i(\lambda) \) is the response function for the filter \( i \). In the numerator, \( R_i(\lambda) \) is integrated only over the limits of the D4000 band of interest (as defined in Table 1), redshifted to the observed frame (hence the redshift of the object, \( z \)). In the denominator, however, the total area of \( R_i(\lambda) \) is computed. Thus, this \( r_i \) factor can be interpreted as the fraction of the NB that is inside the wavelength range of the D4000 band.

Regarding the \( \Delta \lambda_i \) factor in equation (3), it must be a term with wavelength units to make the broader NBs weight more in the calculation of the D4000 bands. The definition we have chosen is simply

\[
\Delta \lambda_i = \int_0^{\lambda_{\text{max}}} d\lambda \ R_i(\lambda).
\]

If the response functions \( R_i(\lambda) \) were ideal top-hat functions with a maximum value of 1, intuitively \( \Delta \lambda_i \) would be the width of the top-hat, and the product \( \Delta \lambda_i F_{v_i} \) would simply correspond to the area of the flux integrated by the NB \( i \). Consequently, equation (3) in this ideal case could be interpreted as a Riemann sum weighted by \( r_i^2 \sigma_i^{-2} \). Hence, equation (3) is the generalization of a weighted Riemann sum to any set of photometric bands with response functions \( R_i(\lambda) \).

The weighting \( r_i^2 \sigma_i^{-2} \) can be interpreted as the contribution of two different error sources: the error of the NB measurement itself \( (\sigma_i) \), and the part of the NB corresponding to the flux integrated outside of the D4000 band of interest. All flux information coming from outside the D4000 band wavelength range can be regarded as noise; therefore an NB which has half of its area outside of the D4000 band should have half the SNR as an NB which is fully inside. This can be expressed in terms of a net error, \( \sigma_{\text{net}} = r_i^{-1} \sigma_i \). Therefore, the weighting is just a standard inverse variance weighting \( (\sigma_{\text{net}}^2) \), but using a \( \sigma_{\text{net}} \) error that accounts for both the intrinsic error of the observation of each band flux, and the undesired flux information outside of the D4000 rest-frame band. Regarding the error of equation (3), it is simply determined by standard error propagation of the errors of the NB fluxes, \( F_{v_i} \).

In Fig. 3 an example of the D4000\(_w\) photometric (with PAUS NBs) and spectroscopic (with the VIPERS spectrum) measurement for a single object is displayed. The selected object is a galaxy with D4000\(_w\) \( \sim 1.7 \), a value high enough to have a spectral break easily distinguishable by eye, and typical for red galaxies. Hence, it can be considered a fairly representative example of a red galaxy in our sample. The apparent magnitude of the object is \( i_{\text{AB}} = 20.71 \). The plot is shown in rest frame, and the values of the D4000\(_w\) band fluxes are determined with equation (2) for the spectroscopic measurement and equation (3) for the photometric measurement, respectively. Besides, the PAUS NB response functions are plotted with a faint grey colour for reference. The relevant data of regarding the D4000 measurement of the object shown in Fig. 3 are displayed in Table 2. It is worth noting that if the PAUS D4000 is computed using the VIPERS...
spec-z, we find a photometric D4000\_w = 1.50 ± 0.14, a result consistent with the PAUS value displayed in Table 2, but not so much with its VIPERS counterpart. The reason for this discrepancy may lie in how the noise in PAUS NBs influences the determination of each photo-z, yielding D4000 values that are more self-consistent than these computed with an external redshift measurement. However, as it is shown in Section 4.1, the actual D4000 error budget for photometric measurements is perfectly accounted for by the errors in PAUS photometry alone.

3.1.2 Synthetic PAUS tests

Before applying it to real PAUS data, the photometric estimator for the D4000 bands defined in equation (3) has been tested against synthetic PAUS NBs, in order to evaluate the potential biases stemming from the estimator alone, without the experimental errors in the PAUS photometry or photo-z, as well as other possible unaccounted systematics.

The synthetic PAUS data set (sPAUS) has been generated by selecting the whole VIPERS sample in W1, and integrating the spectra for all PAUS NBs considered in this work (Fig. 1). These synthetic NBs have been used to determine the D4000 with equation (3), using the respective spectroscopic VIPERS redshifts. Therefore, any deviations from the spectroscopic measurements must be due to our photometric estimation method of the D4000 (which is limited by the spectral resolution of the narrow-band photometry).

A comparison between the VIPERS D4000 and the sPAUS D4000 is shown in Fig. 4, both for D4000\_w and D4000\_h. In both panels, the scatter plot on the left displays the VIPERS D4000 versus the bias of the sPAUS D4000, defined as

\[
\text{Bias (per cent)} = 100 \times \left( \frac{\text{D4000}_{\text{sPAUS}}}{\text{D4000}_{\text{VIPERS}}} - 1 \right),
\]

together with a linear fit and its error (yellow line). The histogram on the right is a histogram of the bias, with its mean value and σ shown in yellow. From these plots, it can be inferred that the sPAUS D4000\_w has a slight mean negative bias of the order of ~1 per cent that seems fairly constant across all the VIPERS D4000 range (the linear fit shows just a small correlation). On the other hand, the sPAUS D4000\_h shows a smaller net bias but also more dispersion, and a clear correlation between bias and spectroscopic D4000 values: the sPAUS D4000\_h is overestimated for bluer galaxies (D4000\_h < 1.4), while it is underestimated for higher values (i.e. for redder galaxies). Moreover, in both cases the total bias distribution in the histograms is fairly close to Gaussian. Some numerical results from this comparison are summarized in Table 3.

Overall, these results show a significant improvement over the preliminary analysis carried out in Stothert et al. (2018), where a simple linear interpolation was used to estimate the photometric D4000 over a similar synthetic catalogue (computed from SDSS spectra instead of VIPERS). In Fig. 10 of Stothert et al. (2018), a comparison akin to Fig. 4 is displayed, but showing only mean biases (without percent) in bins of the VIPERS D4000. While the same trends appear (a clear linear correlation in the D4000\_h), the absolute

| Mean bias (per cent) | σ bias (per cent) | <1 per cent | <2.5 per cent | <5 per cent |
|----------------------|-------------------|-------------|--------------|------------|
| D4000\_w | -1.03 | 2.59 | 41.06 | 77.21 | 95.11 |
| D4000\_h | -0.23 | 3.19 | 32.07 | 67.44 | 92.62 |

Figure 4. The bias of the photometric D4000 measured in the synthetic PAUS NBs, versus VIPERS D4000 values. A linear fit with its error and the mean bias and its σ in the histograms are marked with yellow lines and areas, respectively. Left-hand panel: D4000\_w. Right-hand panel: D4000\_h.

Figure 5. Bias versus z\_apparent for the sPAUS D4000\_w (upper panel) and sPAUS D4000\_h (lower panel). The grey line and shaded area represent the mean bias and its σ for each z\_apparent, while the yellow line and its shaded area are the mean bias and σ for the whole z\_apparent range.
bias values are significantly lower in this work. This is especially noticeable on the low and high D4000 ends, e.g. at D4000 $\sim 2$ Stothert et al. (2018) finds a mean bias of $\sim 7.5$ per cent (dashed line), while in this work the linear fit shows a bias of $\sim 5$ per cent. A similar improvement appears in the D4000$_w$, where they find biases of $\sim 7.5$ per cent and $\sim 2.5$ per cent for the low and high D4000 ends, while Fig. 4 clearly shows smaller mean biases in these regions. Besides, their mean bias of $\sim 2$ per cent for the photometric D4000$_w$ found in Stothert et al. (2018) is reduced by a factor of two.

The reason for this improvement is the integration of the D4000 bands over NBs using an estimator that properly weights the flux fraction in each NB that is relevant to the rest-frame band (equation 3), as well as the observational error. With linear interpolation, NBs mostly outside the rest-frame band can still affect the slope of the interpolated SED (if they are adjacent to NBs ‘fully’ inside the rest-frame band), and thus affect significantly the value of the integrated flux (Fig. 3 may help visualizing this effect).

In addition to the dependence of the sPAUS D4000 bias with the VIPERS D4000, we investigate its correlation with redshift. Given that the D4000 bands are defined in rest frame, the $r_i$ factor in equation (4) and thus the NBs used in the photometric D4000 calculation will be a function of $z$. Therefore, one could expect the redshift regions where the NBs better match the D4000 bands to be less biased and vice-versa (i.e. when the D4000 bands are covered solely by NBs fully inside the band, the measurement should be more accurate). As a consequence, the bias as a function of redshift should exhibit an oscillatory pattern as redshift increases and the
D4000 bands shift to redder wavelengths in observed frame (since there will be an alternation of ‘good’ and ‘bad’ NB matches to the D4000 bands).

It is worth noting that this redshift dependence is just an effect of the relative position of the NBs with respect to the redshifted D4000 bands; any possible effect derived from galaxy evolution or observational systematics dependent on redshift is not considered. Since we are working with synthetic NBs integrated from VIPERS spectra, we can shift the wavelength of the spectra with the following change of variable before NB integration,

$$
\lambda' = \frac{1 + z_{\text{apparent}}}{1 + z} \lambda,
$$

where $z$ is the actual redshift of the object, and $z_{\text{apparent}}$ is the redshift we want the NBs to ‘observe’ the object with. This transformation is akin to shifting the object to the rest frame and then moving it to a different observed frame defined by $z_{\text{apparent}}$. Hence, we can evaluate the redshift bias of any object not only at its real redshift $z$, but at any $z_{\text{apparent}}$, as long as the photometric D4000 can still be measured in the new observed frame.

Consequently, we have evaluated the redshift bias for all objects in the VIPERS W1 sample by computing their NBs in a fine $z_{\text{apparent}}$ grid. The results are shown in Fig. 5. The bias appears as a continuous function because of the small grid size. At each point, the displayed bias and its error are, respectively, the mean bias and its $\sigma$ of all the objects of the VIPERS W1 sample where the photometric D4000 could be measured after applying equation (7).

For the sPAUS D4000$_w$, this redshift bias oscillates $\lesssim 2$ per cent around the mean bias, while in the sPAUS D4000$_n$ the oscillations reach well above 2 per cent. In both cases, the oscillations become significantly smaller as redshift increases, because the larger the redshift, the wider the D4000 bands in observed frame, and the smaller the error induced by NBs partially outside the D4000 band ($0 < r < 1$). Moreover, the oscillations seem fairly stable around the mean bias (yellow lines); hence we can consider the redshift oscillations independent of the D4000 bias (see Fig. 4).

After these tests, we can state that our estimator for NB photometry has a mean bias of $-1.03 \pm 2.59$ per cent and $-0.23 \pm 3.19$ per cent for photometric D4000$_w$ and D4000$_n$, respectively (see Table 3), plus an oscillatory redshift bias due to the relative position of the NBs to the D4000 bands in observed frame. This oscillatory bias adds an extra $\sim 2$ per cent to the D4000$_w$ in the worst cases (yielding a much smaller effect in most of the redshift range), while this increase is also larger for the D4000$_n$. Besides, for the photometric D4000$_w$ there is a clear decreasing linear correlation between bias and D4000 value, while for the photometric D4000$_n$ this correlation is much weaker. In any case, a significant fraction of the sample shows sub-per cent bias (41 per cent and 31 per cent for D4000$_w$ and D4000$_n$, respectively), and the large majority has an absolute bias below 5 per cent (95 per cent and 93 per cent, respectively, see Table 3). Given these results, no bias corrections have been applied to the photometric D4000 measurements, since the errors of the bias are larger than the biases themselves.

Overall, the estimation of the photometric D4000$_w$ through this method seems more reliable, but given the preponderance in the literature of the D4000$_n$, we will consider both for scientific use. Considering that these results are for a synthetic data set generated from spectroscopic data, without the larger photometric and redshift errors from narrow-band photometry, it is safe to assume that the biases of the estimator are negligible, or at least completely subdominant, when compared to the observational errors that we can expect.

### 3.1.3 Redshift bias versus filter FWHM

In addition to analysing the photometric D4000 bias for PAUS NBs, we have further explored the oscillatory redshift bias by repeating the analysis displayed in Fig. 5 for different filter configurations. Instead of using PAUS NBs (Fig. 1), we have defined a set of top-hat filters of different FWHM (between 210 Å and 50 Å approximately) fully covering the same wavelength range as PAUS without overlap between them. For each FWHM, the synthetic NBs have been determined for all objects, as well as the photometric D4000 and its mean redshift bias, by applying the redshift translation defined in equation (7).

By definition, for these ideal filter top-hat sets the filter separation and FWHM are the same; therefore, when we refer to the FWHM of the top-hat filters we also refer to the separation between them. For PAUS however, filter separation is $\sim 100$ Å while their FWHM is $\sim 130$ Å. In a realistic setting, the response function of a filter will never be a top-hat function, and some filter overlap will be inevitable if a reasonably full coverage is to be expected. Regardless, this simple study with top-hat filters may be useful to evaluate if different filter separations may yield less biased D4000 measurements.

Fig. 6 shows the dominant frequency of the redshift bias versus FWHM, determined via fast Fourier transform. Moreover, the mean, maximum, and minimum value of the redshift bias versus FWHM is also evaluated (for the same redshift range as Fig. 5). The mean is the same as the yellow line in Fig. 5, while the maximum and minimum are the highest peak and lowest valley of the oscillatory grey line in Fig. 5, respectively. Both D4000$_w$ and D4000$_n$ show a very similar frequency decrease with FWHM, which seems to asymptotically approach zero as the FWHM of the filters becomes significantly larger than the width of the D4000 bands (Table 1). The mean redshift bias is kept at sub-per cent levels for the whole FWHM range for D4000$_w$, while for D4000$_n$ this is only true at FWHM < 180 Å (after a drastic increase in the absolute bias value at FWHM > 150 Å). Regardless, it seems that the average D4000$_w$ bias can only be improved by reducing the FWHM up to 80 Å, while for D4000$_n$ the current PAUS FWHM already yields an optimal result.

On the other hand, the maximum and minimum redshift biases fluctuate more erratically with the FWHM, as the redshift bias signal is composed of different frequencies (this can be seen in Fig. 5, where the oscillatory signal can be seen to decrease as the FWHM increases).
although for the PAUS case the non-dominant frequency components are not very noticeable). For the case of D4000s, there seems to be a sweet spot around FWHM ~ 100 Å where the redshift bias is contained in the smallest range possible (around ±2 per cent). For D4000s, FWHM < 90 Å provides the smallest redshift bias oscillation, and using even narrower filters does not seem to decrease the bias range. Given that the PAUS filter separation is already 100 Å, the decrease we may expect in D4000 photometric measurement with even narrower bands will be fairly limited. In fact, as we will see in Section 4.1, the noise in band fluxes will be the limiting factor for the photometric D4000 measurement, so the usage of narrower bands may be indifferent or even counterproductive if exposure times are not increased accordingly.

### 3.2 CIGALE estimation

In addition to the direct measurement over PAUS NBs, we have used the SED fitting algorithm Code Investigating GALaxy Emission (CIGALE; Noll et al. 2009; Boquien et al. 2019) to perform a D4000 reconstruction from photometric observations. CIGALE is a physically motivated state-of-the-art Python code for SED fitting based on the principles of the energetic balance between dust-absorbed stellar emission and its re-emission in the infrared (IR). The capabilities of this SED fitting tool have already been verified on PAUS observations (Johnston et al. 2021; Tortorelli et al. 2021). For this work, we have used a delayed star formation history, Bruzual & Charlot (2003) single stellar population models with the initial mass function (IMF) given by Chabrier (2003), a Charlot & Fall (2000) attenuation law and the dust emission models of Dale et al. (2014) to built a grid of models. A detailed description of each module can be found in Malek et al. (2018) and Boquien et al. (2019).

The adopted parameters employed for the SED fitting are listed in Table 4. Thanks to the use of only one free parameter in the Dale et al. (2014) model (i.e. the $\alpha$ slope in $\mathrm{dM}_{\mathrm{dust}} \propto U^{-\alpha} \mathrm{d}U$, where $\mathrm{M}_{\mathrm{dust}}$ and $U$ are dust mass and radiation-field intensity, respectively), the...

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**Figure 10.** Histogram of the D4000s (even rows) and the D4000s (odd rows) for the photometric PAUS direct measurement (left-hand column), PAUS CIGALE reconstruction (middle column), and CFHTLS CIGALE reconstruction (right-hand column). The two upper rows correspond to the bright sample, while the two lower rows display the full sample. All panels include the distribution of the spectroscopic VIPERS D4000 measurement (purple dashed), as well as the distribution of the test data set generated by adding a Gaussian value of mean zero and $\sigma$ equal to its respective non-spectroscopic error to the VIPERS D4000 values (black dashed line).
The SED-fitting procedure reduces the number of parameters to constrain without far-infrared measurements. Those models are then fitted to the galaxy SEDs with the use of a Bayesian-like analysis. For all models, the \( \chi^2 \) and the likelihood (\( \exp - \chi^2/2 \)) for a given observed galaxy are computed, and the value and error of all the parameters returned by CIGALE are the likelihood-weighted mean and standard deviation of all the models, respectively (Boquien et al. 2019).

The quality of the fit is expressed by the reduced \( \chi^2 \) of the best-fitting model, \( \chi^2_r \) (i.e. the \( \chi^2 \) divided by the number of data points). We have applied a quality cut to these fits by removing all objects with a \( \chi^2_r \) larger than median(\( \chi^2_r \)) + 3 (in all CIGALE runs considered). This results in 802 objects being removed from the selected sample described in Section 2.3, before the redshift cuts for D4000\(_w\)/D4000\(_s\) measurement (the galaxy numbers displayed for these redshift ranges already take into account this quality cut).

For this work, SEDs were fitted to galaxies using: (i) PAUS redshift and its corresponding NB photometry (hereafter PAUS CIGALE), and (ii) CFHTLS redshift (Ilbert et al. 2006; Coupon et al. 2009) and its broad-band photometry (hereafter CFHTLS CIGALE). These two data sets allow us to explore the improvement of the D4000 reconstruction accuracy with narrow-band photometry. As one may expect more improvement fitting both NBs and broad-bands, we also run CIGALE with both PAUS NBs and CFHTLS broad-band photometry, but the results were more discrepant with VIPERS spectroscopy than PAUS NBs alone. By adjusting the \( \chi^2 \) computation procedure (Eriksen et al. 2019), the combination of PAUS and CFHTLS photometry should yield better results than PAUS alone; however, any modifications to CIGALE are left as future work.

### 4.1 SNR

We have computed both D4000\(_w\) and D4000\(_s\) for the sample specified in Section 2.3, and examined the signal-to-noise ratio (SNR) for each method. Fig. 7 displays the SNR distribution for the D4000 estimations with each method (PAUS direct, CIGALE runs, and VIPERS spectroscopy), together with the detection threshold we have imposed at SNR > 3. This SNR limit has been chosen because it is commonly used in astrophysics as a threshold for source detection in imaging or spectral features, whether it is in the visible spectrum (e.g. Cutri et al. 2003; Sánchez et al. 2012), X-ray (e.g. Bulbul et al. 2014), or infrared wavelength range (e.g. Labbé et al. 2013). The distribution of VIPERS SNR appears as a dashed line; throughout this paper we will use the convention of representing data based on photometric redshifts with solid lines and spectroscopic redshift data with dashed lines.

As shown in Fig. 7, the SNR from the CIGALE D4000 reconstruction is significantly larger than the one from PAUS direct. The D4000 values from both CIGALE runs have a mean SNR of ~20, being slightly larger for the CFHTLS case. In both CIGALE runs, all objects in the selected sample are well above the detection threshold, while in the PAUS direct case only 85.01 percent (65.87 percent) of D4000\(_w\) (D4000\(_s\)) measurements have SNR > 3. The mean SNR from the PAUS direct method is ~5 for the D4000\(_w\) and ~4 for the D4000\(_s\). Regarding the spectroscopic VIPERS D4000, we find a mean SNR of ~35 for the D4000\(_w\) and ~20 for the D4000\(_s\). The larger SNR in the D4000\(_w\) for both photometric and spectroscopic measurements is simply due to the larger wavelength range of the D4000 bands for the D4000\(_w\) definition: if flux is integrated over a larger span, higher SNR is to be expected. The CIGALE D4000 reconstruction, however, seems insensitive to this effect, given that it uses all available flux information to fit a model that predicts a D4000 value. This results in an SNR that is closer to spectroscopy for D4000\(_w\).

It is reasonable to expect a higher SNR from SED fitting than PAUS direct measurements, given that the spectral information provided to CIGALE spans a much larger wavelength window. CIGALE uses all reference broad-bands or PAUS NBs, while in the PAUS direct measurement just the NBs that may contain part of the D4000

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**Table 4.** The input parameters used for SED fitting with CIGALE.

| Parameter | Delayed star formation history |
|-----------|--------------------------------|
| e-folding time of the main stellar population model (Myr) | 100, 300, 500, 1000, 2000, 5000 |
| Age (Myr) | 300, 500, 1000, 1500, 2500, 4500, 6000 |
| Initial mass function | Chabrier (2003) |
| Metallicity \([Z\odot]\) | 0.02 |
| Age of separation between young and old stellar populations (Myr) | 10 |
| Dust attenuation law (Charlot & Fall 2000) | 0.001, 0.1, 0.4, 0.7, 1.0, 1.5, 2.0, 2.5 |
| V-band attenuation \((A_V)\) in the interstellar medium (ISM) | –0.7 |
| Power-law slopes of the attenuation in the birth clouds (BC) and ISM | 0.8 |
| AGN fraction | 0.0, 0.1, 0.3 |
| Power-law slope \(dU/dM\) \((L^\odot)\) | 2.0 |
| Ionization parameter | \(10^{-2}\) |
| Escape fraction of Lyman continuum photons | 0.0 |
| Absorption fraction of Lyman continuum photons | 0.0 |
bands are taken into account. Moreover, the SED fitting relies on underlying galaxy models, which already places constraints on the possible values of the spectral features, thus constraining the potential dispersion of D4000 values. Nevertheless, these factors alone cannot account for the extreme SNR boost in the CIGALE reconstruction. It is worth noting that CIGALE only uses the observational errors as weights for the \( \chi^2 \) computation, without propagating them directly into the error budget; the error of the output values (in this case, the D4000) is the likelihood-weighted standard deviation of all fitted models (Boquien et al. 2019). In other words, the noise levels of the photometric fluxes will determine how much these fluxes are weighted in the fit relative to each other, but do not have a direct relationship to the error of the parameter estimated with CIGALE. Therefore, if the PDF of the fit is concentrated in a small probability volume (either because the models do not properly reproduce the diversity of the sample, or because the \( \chi^2 \) computation tends over weight specific input data), the error of the output values will be low, even if the input data is noisy. This may result in an error underestimation, and thus explain the high SNR values of CIGALE. The possibility of an error underestimation due to a limited model grid or unaccounted photometric errors is largely discussed in Noll et al. (2009). However, any modifications or improvements to the SED-fitting code are out of the scope of this work.

The Gaussianity of the errors associated to these D4000 estimations is evaluated in Fig. 8 and Table 5. There, we display the distribution of the difference between the D4000 estimations and the true value (i.e. VIPERS spectroscopic measurements), divided by the estimated error. If the errors are properly estimated, these histograms should follow Gaussian distributions of \( \mu = 0 \) and \( \sigma = 1 \). Errors from PAUS direct photometric measurements are realistic, as their histograms in Fig. 8 have \( \sigma \sim 1 \) (underestimated by \( \sim 10 \) per cent for D4000\(_a\)) and overestimated by \( \sim 5 \) per cent for D4000\(_b\), see Table 5). Both CIGALE methods show clear error underestimations, with \( \sigma \sim 1.75 \) for CFHTLS CIGALE, and \( \sigma \sim 1.5 \) for PAUS CIGALE, which is akin to an error underestimation of 75 per cent and 50 per cent, respectively. Hence, the SNR values derived from the CIGALE reconstructions in Fig. 7 should be divided at least by a factor equal to their \( \sigma \) in Table 5 to be considered realistic. Despite the PAUS direct D4000 measurement providing a realistic error estimate consistent with a Gaussian behaviour, and the CIGALE D4000 reconstruction overestimating its SNR, it is worth noting that by construction, CIGALE still provides less dispersion. Due to the limited grid of SED models, the reconstructed CIGALE D4000 will always have a realistic value within the range allowed by the models, regardless of how faint the observed object is.

Given the significant percentage of objects with SNR < 3 in PAUS direct, we have evaluated up to which magnitude PAUS direct allows an individual photometric D4000 measurement of all galaxies above the established SNR threshold (SNR > 3). In Fig. 9, the percentage of objects with SNR > 3 in apparent magnitude bins is shown. Both photometric D4000\(_a\) and D4000\(_b\) have SNR > 3 for almost all objects up to \( i_{AB} < 21 \) (99.88 per cent and 98.97 per cent for D4000\(_a\) and D4000\(_b\), respectively), with the percentage of objects with measurable D4000 decreasing above this threshold. Consequently, we define an SNR > 3 D4000 subsample by applying an additional \( i_{AB} < 21 \) cut to the sample defined in Section 2.3; this leaves a total of 2254 galaxies. We will refer to this \( i_{AB} < 21 \) sample as bright sample henceforth, while the original sample defined in Section 2.3 will be referred to as full sample. We will use this bright sample to evaluate how well PAUS direct performs when all objects have a SNR acceptable for individual measurements, while the full sample will be employed to evaluate how well can we recover mean sample trends despite the low SNR.

The advantage of this bright sample is that it preserves PAUS magnitude completeness (which large spectroscopic surveys do not achieve, due to target selection) while ensuring that the PAUS direct individual photometric measurements have high enough SNR. For the bright sample, the average PAUS direct SNR is 11.42 for D4000\(_a\) and 8.63 for D4000\(_b\). At \( i_{AB} = 23 \), the magnitude limit for the PAUS catalogue, only \( \sim 30 \) per cent of the objects have a D4000\(_b\) above detection threshold for the photometric measurements, while for D4000\(_a\), this percentage decreases to just \( \sim 12 \) per cent. Hence, with the narrow-band configuration of PAUS we can achieve D4000 individual measurements with SNR > 3 for a magnitude-completeness threshold two magnitudes brighter than the magnitude limit of the full catalogue. It is reasonable to expect similar performance with comparable narrow-band surveys, such as J-PAS (Benitez et al. 2014).

### 4.2 Distribution and bias

Once we have examined the SNR and error of the D4000 estimation methods, we study the distribution of the D4000 values themselves. Fig. 10 shows a histogram of D4000 values for each method, and for both the bright and full samples. In addition to this, we have also generated test data sets for PAUS direct, PAUS CIGALE, and CFHTLS CIGALE D4000 by taking the spectroscopic VIPERS D4000 measurements for each object, and adding to them a value drawn from a Gaussian distribution of mean zero and \( \sigma \) equal to the error of its respective non-spectroscopic estimation. For the CIGALE D4000 reconstructions, we use the error provided by CIGALE as it is, without any correction to mitigate the error underestimation shown in Fig. 8 and Table 5.

If the D4000 estimations from these non-spectroscopic methods are unbiased and with an error estimation that behaves like a Gaussian, the D4000 distribution from the test data set generated by adding these artificial errors to VIPERS D4000 should be the same as the D4000 distribution from its respective non-spectroscopic method. We know from Fig. 8 that the error in the photometric PAUS direct measurement is reasonably close to Gaussian, and that for the CIGALE reconstructions it is clearly underestimated; these test data sets allow us to check if the differences between the VIPERS D4000 distribution and the non-spectroscopic D4000 distributions can be accounted for solely by the non-spectroscopic error. The test data sets are represented by dashed black lines.

The spectroscopic D4000 distributions in Fig. 10 (purple dashed line) exhibit a weak bimodality, with a peak centred on the blue cloud around 1.3 and 1.2 for D4000\(_a\) and D4000\(_b\), respectively. This blue peak is far stronger for the full sample: as blue galaxies are fainter on average, by imposing the \( i_{AB} < 21 \) cut most of the removed galaxies belong to the blue cloud. On the other hand, the red sequence appears as a plateau in the much longer tail of the

### Table 5: Parameters of the Gaussian fits to all error distributions displayed in Fig. 8.

| Method    | Parameter | Mean (\( \mu \)) | Standard Deviation (\( \sigma \)) |
|-----------|-----------|------------------|-----------------------------------|
| PAUS direct | D4000\(_a\) | \(-0.10 \pm 0.03\) | \(1.10 \pm 0.02\) |
| PAUS direct | D4000\(_b\) | \(0.06 \pm 0.02\) | \(0.96 \pm 0.02\) |
| PAUS CIGALE | D4000\(_a\) | \(0.38 \pm 0.02\) | \(1.51 \pm 0.02\) |
| PAUS CIGALE | D4000\(_b\) | \(0.27 \pm 0.02\) | \(1.53 \pm 0.02\) |
| CFHTLS CIGALE | D4000\(_a\) | \(-0.44 \pm 0.02\) | \(1.72 \pm 0.02\) |
| CFHTLS CIGALE | D4000\(_b\) | \(-0.57 \pm 0.02\) | \(1.77 \pm 0.02\) |
The bimodality in the D4000 distribution has been extensively demonstrated in the literature (see e.g. Kauffmann et al. 2003b; Haines et al. 2017) and it has been shown to become less pronounced as redshift increases (e.g. Haines et al. 2017, fig. 1). Therefore, the weak bimodality in our sample at z > 0.5 is to be expected.

In order to quantify the agreement between the D4000 distribution of the different methods and the VIPERS D4000, we have performed the Epps–Singleton (ES) test (Epps & Singleton 1986) on all cases displayed in Fig. 10, comparing the D4000 distributions against the noiseless spectroscopic VIPERS D4000 measurements, and the VIPERS test data sets generated with the non-spectroscopic noise (black lines in Fig. 10). The ES test has been chosen instead of other popular alternatives such as the Kolmogorov–Smirnov (Massey 1951), Cramer-von Mises (Anderson 1962), or Anderson–Darling (Anderson & Darling 1952) tests because it was the least dependant on the D4000 range where the empirical distribution function (EDF) was determined. This result is expected, given that the ES test has been specifically developed to provide enough power under a great variability of EDF parameters, such as location, scale, or even family of distributions (see Epps & Singleton 1986), and in our case the tails of the D4000 distribution greatly vary between D4000 methods. For the PAUS direct measurement, the tails are significantly larger than VIPERS due to the lower SNR, and for both CIGALE reconstructions the tails end abruptly given the limitations of the SED models (see Fig. 10).

The D4000 EDF has been computed in all cases with a uniform binning between D4000 = 0.7 and D4000 = 2.5, with a bin width \( \Delta D4000 = 0.02 \); this is the same binning as the histograms displayed in Fig. 10. The \( p \) values of the ES tests are shown in Table 6; these are the probabilities of the estimated D4000 values coming from the same distribution as the noiseless spectroscopic VIPERS D4000 (or VIPERS D4000 with non-spectroscopic noise, in parentheses).

The distribution of the PAUS direct D4000 (left-hand column in Fig. 10) shows the weakest bimodality of all the non-spectroscopic methods, with a much smaller peak in the blue cloud, and a less pronounced difference between the high and low D4000 tails. This effect is caused by the noise in the photometric D4000 measurements, and is far more noticeable for the full sample due to the much lower average SNR. In fact, the \( p \) values for the bright sample when compared to spectroscopic VIPERS measurements are far from negligible (up to 0.5 in D4000), see Table 6) while for the full sample they tend to zero. However, for the PAUS direct measurements the black line representing the spectroscopic VIPERS D4000 with PAUS direct Gaussian errors shows an excellent agreement (\( p \) value ~ 0.9 and \( p \) value ~ 1 for the bright and full samples, respectively), with the only noticeable difference being a shift of the distribution peak of \( \sim 0.1 \) towards lower values. This is consistent to a certain extent with the negative bias of the photometric PAUS direct estimator shown in Section 3.1.2. Nevertheless, this similarity between distributions is a clear indicator that the error in PAUS direct is close to Gaussian.

Regarding the PAUS CIGALE reconstruction, we can see in Fig. 10 (centre column) that the D4000 distribution is much less affected by noise than in the PAUS direct case. This difference is far more noticeable in the full sample, where the PAUS direct distribution is heavily smoothed by noise, but the PAUS CIGALE distribution follows the VIPERS D4000 distribution far more closely (\( p \) value = 0.84 and \( p \) value = 0.52 for D4000, and D4000, respectively). When we add the CIGALE D4000 reconstruction error to the VIPERS D4000 distribution, the differences between the CIGALE and VIPERS D4000 distributions increase (a drop of \( \sim 0.3 \) in \( p \) values), since the Gaussian error further smooths the VIPERS D4000 distribution, but the PAUS CIGALE distributions are actually ‘sharper’. This ‘sharpening’ of the PAUS CIGALE distributions is an artefact of the SED-fitting methodology; the reconstructed D4000 values tend to cluster around the underlying SED templates, resulting in stronger peaks around the most representative red/blue galaxies than in the real sample (i.e. artificial bimodality). The artificial bimodality is far less evident with the D4000 reconstruction, which may be due to the smaller difference in D4000 values between SED templates due to the larger continuum range, or the larger dependence on dust extinction smoothing this difference. In the case of the CFHTLS CIGALE distribution (Fig. 10, right-hand column), the artificial bimodality is far more exaggerated, thus resulting in a much larger discrepancy with spectroscopy (lower \( p \) values all across Table 6).

In addition to the D4000 distributions, we also study the bias of each D4000 estimation method versus the spectroscopic D4000 value (see Fig. 11, which is the observational counterpart of Fig. 4). In Fig. 11, each panel displays a scatter plot with the VIPERS D4000 measurements in the x-axis, and the bias of the respective D4000 estimation method in the y-axis. This bias is determined with equation (6); but this time with the D4000 estimated from actual PAUS observations (either with direct photometric measurements or CIGALE reconstructions). Both the bright and full sample are displayed, following the same panel arrangement as in Fig. 10. Overall, the biases present a much larger scatter in Fig. 11 than in Fig. 4, especially in the full sample, but this is due simply to the photometric noise, which simply does not exist in Fig. 4. Hence, our analysis in this subsection will be focused on the median bias and the average linear correlation between bias and the spectroscopic D4000 value.

We have performed linear fits to all data sets to evaluate the average correlation between the bias and the VIPERS D4000; all fits are unweighted except the PAUS direct measurement, where an SNR \( ^2 \) weighting has been applied. The reason for this weighting is that the high noise of the PAUS direct D4000 resulted in a linear fit with large error regions, which provided close to no meaningful information about the actual trend. An SNR \( ^2 \) weighting has been chosen instead of \( \sigma ^2 \) to mitigate the dependence of the error with the D4000 value: when weighting objects with different D4000, the SNR is a more accurate estimator of the amount of information per measurement. We have tested this weighting by repeating the SNR \( ^2 \)-weighted linear fit on the CIGALE D4000 reconstructions, but using PAUS direct measurement SNR: the resulting linear fit was compatible with its unweighted counterpart (the error regions of the fits fully overlapped). Hence, we can affirm that the SNR \( ^2 \) weighting does not significantly distort the correlation between bias and VIPERS D4000. It is worth noting that such weighting can not be performed with the CIGALE errors, as there is a far more
significant dependence of errors versus D4000 (almost an order of magnitude, with errors increasing with D4000 value, reaching a maximum around D4000 ~ 1.8). This error–D4000 correlation in the CIGALE D4000 reconstructions distorts the trend found in the unweighted linear fit.

The PAUS direct measurement (left-hand column) displays a mean negative bias of the order of few per cent, with the linear fit showing a negative correlation between the D4000 value and bias. This negative correlation comes from the low spectral resolution of the photometric data used. Since PAUS NBs have a FWHM of the same order of magnitude as the D4000 bands, the red D4000 band may contain some flux information from the slope of the spectral break itself, which would bias its measured flux towards higher values. The opposite may happen with the blue D4000 band, where including flux from the slope of the spectral break will decrease the measured flux (Fig. 3 may help better visualizing this effect). The result of these two effects is an underestimation of the D4000 that becomes more pronounced as the D4000 value increases (see equation 1), a trend consistent with the linear fit to the bias displayed in Fig. 11.

Following with the PAUS CIGALE comparison, the D4000ₙ measurement shows a weaker bias correlation than D4000ₑ (with the slope of the linear fit being approximately half as steep), but a larger absolute value of the median bias (by 2–3 per cent). Given that a fixed bias offset is far simpler to account for and correct than a linear trend depending on the true D4000 values, we consider the PAUS direct measurement of the D4000ₑ to behave better than the D4000ₙ. When comparing the full sample against the bright sample, we see that the bias slope is slightly steeper, and that the absolute value of the median bias is actually smaller by 2–3 per cent. Both of these effects can be accounted for by the low SNR of faint objects in the PAUS sample, given that the D4000 is the ratio of two fluxes. If the flux in the denominator (the blue band, see equation 1) is noisy enough to be close to zero, it will result in an unrealistically high D4000 value (albeit with very low SNR). If we apply a very simple cut to these
high outliers (i.e. PAUS direct D4000 < 5), both the median bias and the slope of the full sample become significantly closer to the bright sample results.

Regarding the PAUS CIGALE D4000 reconstruction (centre column), the median bias has an absolute value of the order of ~1 per cent; however, it is not strictly negative as in PAUS direct. There also appears a negative correlation between VIPERS D4000 and bias, with a slope roughly similar to PAUS direct (slightly steeper for D4000 n, and a bit less pronounced for D4000 e). In this case, the negative slope is mostly caused by the artificial limits that the SED fitting imposes on the D4000 values; it can be clearly appreciated in the centre column of Fig. 11 that the distribution of the data points are clearly cut by two diagonal lines at the low and high D4000 ends. The objects with high VIPERS D4000 values display a strong negative bias (e.g. at VIPERS D4000 n > 1.9), and the objects with low VIPERS D4000 values have a pronounced positive bias (e.g. at VIPERS D4000 e < 1.0).

Comparing the CFHTLS CIGALE results (left-hand column) to PAUS CFHTLS, we see that for CFHTLS the correlation slopes are somewhat less pronounced, but there is a negative median bias of the order of ~2–3 per cent, similarly to PAUS direct. Moreover, there is a strong clustering of data points in a stripe pattern, due to the tendency of the SED-fitting methodology to infer specific D4000 values. One possible explanation for these stripes is the finite number of SED templates in the SED-fitting code, with the D4000 reconstructions tending to cluster around the most representative SEDs for red and blue galaxies (explaining also the stronger bimodality). This stripe pattern is much less noticeable in PAUS CFHTLS, which is in accordance with the smoother D4000 distribution in Fig. 10.

Therefore, we can conclude from this section that the PAUS direct D4000 measurement is fully consistent with the VIPERS D4000 plus a Gaussian noise, which is already estimated by simple propagation of the photometric errors. In the bright sample (where PAUS direct D4000 SNR > 3), the performance is similar to CIGALE SED fitting with PAUS photometry, with the advantage of being completely model-independent (and requiring negligible computational cost). In the full sample, where a significant fraction of the PAUS direct D4000 measurements have SNR < 3, the results can be kept consistent if measures are taken to account for this noise (i.e. simulating the noise in the VIPERS D4000 in Fig. 10, or using SNR 2 or median-based statistics in Fig. 11). Regarding the CIGALE results, we see that PAUS CIGALE performs equally well in both the bright and full samples, and significantly outclasses the CFHTLS CIGALE reconstruction both in error estimation (Table 5) and D4000 distribution (Fig. 10), while providing arguably less biased measurements (Fig. 11). This is a clear proof of narrow-band photometry providing a net benefit to SED-fitting analysis.

Henceforth, we will not take into account the CFHTLS CIGALE D4000 in our comparisons, as we consider proven that the PAUS CIGALE D4000 outperforms it. Moreover, when we evaluate the evolution of average sample trends, we will use the full sample, as the low SNR of the PAUS direct D4000 should not be an impediment to retrieve results compatible with spectroscopy. On the other hand, when we analyse properties related to individual galaxies (e.g. galaxy classification), we will consider both the bright and full samples, in order to evaluate how well the PAUS direct photometric measurement performs in a sample where all objects have SNR > 3.

5 D4000 AND GALAXY PROPERTIES

Here, we examine how the D4000 estimations vary across different observational parameters. For all the results displayed in this section, we have applied additional cuts to our galaxy sample: all objects with SNR < 3 in the VIPERS D4000 (either D4000 n or D4000 e) have been removed, as well as spectroscopic outliers. We have considered outliers all objects outside of the range 0.7 < VIPERS D4000 < 2.5, again, for both D4000 definitions. These cuts removed an extra 46 objects from the sample defined in Section 2.3, resulting in 17 195 galaxies.

5.1 D4000 versus redshift

The first parameter we have considered is redshift: Fig. 12 displays scatter plots of the different D4000 measurements versus redshift, for the full sample. PAUS data (PAUS direct measurements and PAUS CIGALE D4000 reconstructions) have PAUS photo-z in the x-axis, and for VIPERS D4000 measurements its respective spec-z is used. In the panels using PAUS photo-z a sharp stripe of data points appears around z ~ 0.7, but is absent from VIPERS spec-z. This is due to the large horizontal stripes in Fig. A1 around that same redshift. While this is an artefact of the current photo-z computations, it does not affect the average D4000 properties, as it is seen throughout this section. Moreover, such artefacts should be mitigated in future data releases.

In order to quantify the D4000 evolution with redshift, we perform linear fits to the data points, using an SNR 2 weighting for PAUS direct and no weighting for PAUS CIGALE and VIPERS (following the reasoning described in Fig. 11). We have confirmed that the SNR 2 weighting does not distort the trend by applying the same weighting (with PAUS direct SNR) to the VIPERS D4000, and retrieving linear fits compatible with the unweighted results. For all cases, we find a decreasing trend of the D4000 with redshift: for the PAUS direct measurement, this trend is in agreement with 1σ with spectroscopy. However, for PAUS CIGALE the slope of the linear fit is approximately twice as steep as its spectroscopic counterpart (despite PAUS CIGALE mimicking better the true VIPERS D4000 distribution). The reason for this exaggerated trend in the PAUS CIGALE reconstruction might lie in the smaller wavelength range of PAUS NBs compared to CFHTLS broad-bands: at a given redshift, the wavelength window to perform the SED fitting is smaller, even if the spectral resolution is higher. This may result especially problematic in the edges of our redshift range, where the D4000 bands are close to the wavelength limits. Moreover, undesired biases at these redshift limits would affect more the slope of the linear fit than at intermediate redshift values.

At higher redshift stellar populations are younger, and thus a decrease of the average D4000 is to be expected, given that it is a proxy for stellar ages. Despite the selection effects at different redshifts of a magnitude-limited sample, this decreasing trend is noticeable in Fig. 12. This trend is also widely documented in the literature. In fact, with the same VIPERS data used in this work (Scodell et al. 2018), a decreasing trend of D4000 e versus redshift can be seen in Siudek et al. (2017), fig. 7, and Haines et al. (2017), table 1 (far more examples can be found in literature at different redshifts, e.g. Morelso et al. 2012; Joshi et al. 2019; Borghi et al. 2021).

5.2 Galaxy classification

The classification of galaxies into different categories has been one of the main topics of galaxy study ever since the very beginnings of observational cosmology, with the first example being the morphological classification into early-type (elliptical) and late-type (spiral) galaxies (Hubble 1926; de Vaucouleurs 1959). More sophisticated
marginalized morphological classifications have been developed (e.g. Strateva et al. 2001; Driver et al. 2006; Mignoli et al. 2009; Krywult et al. 2017), as well as classifications based on other observables: colours (e.g. Arnouts et al. 2013; Taylor et al. 2015), spectral features (e.g. Kauffmann et al. 2003a), or a combination of them (e.g. Siudek et al. 2018; Turner et al. 2021), including the D4000. It is natural for the simplest classification models to separate galaxies into two different types that mimic the bimodality in their parameters. In this subsection, we will evaluate how well a D4000-based classification with the accuracy of the different methods can separate the blue and red galaxy populations, as defined by a fiducial classification.

The fiducial classification used in this work is presented in Siudek et al. (2018); it has been carried out over the entire VIPERS data release (Scodeggio et al. 2018) with a Fisher Expectation–Maximization unsupervised algorithm, using as entry parameters 12 different rest-frame magnitudes (normalized to $i_{AB}$) and spectroscopic redshift. The algorithm have distinguished 11 different galaxy classes, with an extra class of broad-line AGNs that has been omitted for this work. These classes are defined in order: class 1 corresponds to the reddest galaxies, and class 11 to the bluest ones. For the binary red/blue classification used in this work, we have considered all objects of class 6 and below to be red, and 7 and above to be blue.

This separation between red/blue galaxies is also visible on the NUV $− r$ versus $r − K$ diagram (NUV$rK$ Arnouts et al. 2013; Moutard et al. 2018). In Fig. 13, we show the NUV$rK$ colour diagram for the bright sample ($i_{AB} < 21$), with the data points colour-coded with their PAUS direct D4000$_n$ measurements, and the NUV $− r$ and $r − K$ colours being extracted from the PAUS CIGALE run. The average positions in the diagram of the PAUS red and blue galaxies, according to our fiducial classification, are also displayed and are coherent with the aforementioned references, as well as with the average positions of galaxy classes in the NUV$rK$ colour diagrams (see fig. 1 in Siudek et al. 2018).

This fiducial galaxy classification is used to test the performance of a single D4000 cut (hereafter D4000$_{cut}$) in order to distinguish...
between red and blue galaxies, with a galaxy being blue if $D4000 < D4000_{\text{cut}}$, and red otherwise. For each $D4000$ estimation method, we have tested $D4000_{\text{cut}}$ values in a physically meaningful range that comprises all but the reddest/bluest galaxies ($1.1 < D4000_{\text{cut}} < 1.9$), and determined the percentage of galaxies correctly classified as red/blue with each cut. This $D4000_{\text{cut}}$ range has been chosen because it is unreasonable to evaluate $D4000_{\text{cut}} = 1$ or $D4000_{\text{cut}} = 2$ as thresholds to separate between red and blue galaxies, since these values correspond to highly star-forming and very quenched galaxies, respectively.

The results are given in Fig. 14 and Table 7. In order to check the consistency of these results with the literature, we also evaluate the percentage of correctly classified galaxies using the cut of $D4000_{\text{cut}} = 1.55$ as defined in Kauffmann et al. (2003b) (and other works, e.g. Haines et al. 2017). Since this cut is defined for the $D4000_{\alpha}$, we do not perform this comparison for the $D4000_{\alpha}$. This $D4000_{\text{cut}} = 1.55$ value is significantly higher than the $D4000_{\text{cut}} = 1.4$ we find for the VIPERS $D4000$ measurements in the full sample. This discrepancy may simply arise from the different classification criteria adopted, but it is remarkable that the difference between the $D4000_{\text{cut}}$ derived from the fiducial classification (Siudek et al. 2018) and the $D4000_{\text{cut}} = 1.55$ proposed in Kauffmann et al. (2003b) is smaller than the difference between the $D4000_{\alpha}$ and $D4000_{\alpha}$ cuts derived from the fiducial classification alone. Hence, the $D4000$ definition may play a larger role in the value of the optimal value of the $D4000$ cut than the classification criteria themselves.

The PAUS direct $D4000$ measurement achieves the lowest percentage of correct classification due to the photometric noise, with at most 81 per cent of objects correctly classified (for the bright sample using the $D4000_{\alpha}$). This is also the instance where its effectiveness at classifying becomes closest to the VIPERS $D4000$ spectroscopic measurement (85 per cent). Using the full sample degrades the percentage of correct classification of the PAUS direct cut, given that there is a significant fraction of objects with SNR < 3. This also translates in a flattening of the curve in Fig. 14; in fact, for the $D4000$, this curve is almost horizontal after the optimal $D4000_{\text{cut}}$ value, which shows that varying $D4000_{\text{cut}}$ barely has any effect on the reliability of the classification, and thus the correlation between galaxy type and $D4000$ value is weak.

PAUS CIGALE $D4000$ outperforms the VIPERS spectroscopic measurement in its respective optimal $D4000_{\text{cut}}$ (at least 89 per cent of correctly classified objects). This is reasonable, given the model-driven bimodality/clustering that appears in the CIGALE $D4000$ reconstructed values, stemming from the SED templates (Fig. 10). Moreover, the PAUS CIGALE $D4000_{\text{cut}}$ is the closest to spectroscopy in almost all cases (within ±0.02, except for $D4000_{\text{cut}}$ for the bright sample). This shows that the PAUS CIGALE $D4000$ provides realistic $D4000$ values in the green valley between the red and blue populations, where the $D4000$ cut is placed. PAUS direct also yields values of $D4000_{\text{cut}}$ close to the VIPERS results (within ±0.02), but only for the bright sample (see Table 7); the mean SNR is too low in the full sample to provide a reliable galaxy classification.

It is worth noting that, for the case of $D4000_{\text{cut}} = 1.55$ (values in brackets in Table 7), the difference in the percentage of correctly classified galaxies between the VIPERS $D4000$ spectroscopic measurement and the CIGALE $D4000$ reconstruction is significantly smaller (within ∼3 per cent maximum) than for our fiducial classification. This result is reasonable: if we use a $D4000$ cut already defined in the literature for spectroscopic $D4000$ measurements, the SED-fitting approach of CIGALE does not outperform the spectroscopic VIPERS $D4000$ nearly as much.

In order to check if these results were consistent with a different criterion for red/blue galaxy classification, we have recomputed the optimal $D4000$ threshold and its percentage of correctly classified objects, as in Fig. 14, but using as a fiducial classification the spectroscopic VIPERS $D4000_{\text{cut}}$ shown in Table 7 for the bright sample. The results are largely the same (similar maximum percentage and optimal $D4000$ thresholds), and thus we can state that the performance of the $D4000$ estimations for galaxy classification is not largely dependent on the fiducial classification itself.

5.3 $D4000$ versus stellar mass and SFR

The dependence of the $D4000$ versus stellar mass and SFR (Figs 15 and 16, respectively) is evaluated for the full sample. For the $D4000$–mass relation, we separate the sample into red and blue red galaxies based on the fiducial classification (see Section 5.2). The errorbars are the median absolute deviation (MAD) of the $D4000$ estimations. The MAD has been multiplied by its respective factor (∼1.4826) so it is comparable to the standard deviation, and has not been divided by the square root of the number of objects per bin, so it accounts for the intrinsic variance of $D4000$ values inside a bin (therefore, it should remain constant for a number of objects large enough). We have used the median and the MAD to evaluate the dependence of the $D4000$ due to its insensitivity to outliers.

In order to quantify the $D4000$–mass relation in Fig. 15 we have performed linear fits to the median values of the stellar mass bins; their slopes are specified in Table 8. The PAUS direct $D4000$ mass slope shows a good agreement with VIPERS for blue galaxies, with the slope values agreeing well within their error intervals derived from the fit. On the other hand, the slopes of the red population only agree for $D4000_{\alpha}$; for $D4000_{\alpha}$ PAUS direct significantly
Table 7. D4000cut that ensures the maximum percentage of correctly classified red/blue galaxies, for all cases represented in Fig. 14. The percentage of correctly classified galaxies with this optimal D4000cut is displayed in parentheses, while the same percentage for the standard D4000cut = 1.55 is shown in brackets (only for D4000b).

|                | PAUS direct | PAUS CIGALE | VIPERS |
|----------------|-------------|-------------|--------|
| Bright sample (iAB < 21) |             |             |        |
| D4000cutw      | 1.54 (81.45 per cent) | 1.57 (89.86 per cent) | 1.57 (85.36 per cent) |
| D4000cutn      | 1.39 (76.36 per cent) | 1.38 (88.99 per cent) | 1.42 (84.10 per cent) |
| Full sample (iAB < 23) |             |             |        |
| D4000cutw      | 1.65 (75.67 per cent) | 1.58 (90.61 per cent) | 1.57 (87.20 per cent) |
| D4000cutn      | 1.53 (69.13 per cent) | 1.38 (89.98 per cent) | 1.40 (85.13 per cent) |

Figure 15. Median D4000 in uniform bins of log M∗, for PAUS direct measurement (orange errorbars), PAUS CIGALE reconstruction (cyan), and VIPERS measurement (purple). The fits have been computed to the median values displayed in the plot. The errorbars represent the MAD of the D4000 of the sample. The black dotted errorbars in VIPERS show the MAD of the sample after adding to each object a Gaussian error with the same σ as its respective PAUS direct measurement (as in Fig. 10). The errorbars for each measurement method have been offset in the x-axis for clarity, but in all cases the same binning has been used (as indicated by the grey vertical lines). Middle and lower rows are for D4000w and D4000n respectively, while left-hand and right-hand columns are for blue and red galaxies, respectively, according to the classification of Siudek et al. (2018). The upper row displays a histogram of the number of objects for the respective population.

underestimates the slope. This is a consequence of the negative D4000–bias correlation shown in Fig. 11; for low (high) true D4000 values, the PAUS direct measurements overestimate (underestimate) the D4000. The MAD of the test data set (VIPERS with PAUS direct σ) in Fig. 15 is also very similar to the MAD of PAUS direct D4000 for most stellar mass bins (especially for D4000b). This agreement reinforces the already discussed fact that the PAUS direct dispersion is fairly coherent with a Gaussian realisation of the error.

The PAUS CIGALE D4000 slopes agree with the VIPERS ones within their error intervals for all cases, with D4000w and D4000n performing equally well (Table 8). However, the CIGALE D4000 reconstruction shows in the high-mass end of the red sample and the low-mass end of the blue sample a smaller MAD than the VIPERS D4000 (by a factor of ~3 and ~2, respectively). This underestimation is a clear sign that, despite correctly determining the average spectral features of a given population, the SED templates may not fully reproduce the actual diversity of the sample in the most ‘extreme’ regimes: the bluest and least massive star-forming galaxies, and the most massive and quenched red galaxies.

We can compare the slopes from Table 8 with previous results from the literature. In Haines et al. (2017), the D4000b–mass relation is derived for red and blue galaxies separated by a D4000cutn = 1.55, using the same VIPERS data release (Scodellaro et al. 2018). They reported a blue D4000b–mass slope between 0.12 and 0.14, invariant with redshift; a result reasonably close to the slope of 0.11 we report in Table 8. Regarding the red slope, Haines et al. (2017) derived a
value of \( \sim 0.23 \) at 0.5 < \( z \) < 1.0, a slope significantly less steep that the value of 0.33 we find. Another work that analyses the D4000\(_v\)–mass relation for the red sequence is Siudek et al. (2017), where an evolving \( U - V \) colour cut is used to select red galaxies. Here, a slope of 0.164 \( \pm \) 0.031 is found for the same VIPERS data release at 0.40 < \( z \) < 1.00; a value in even larger disagreement with our results.

The reason behind these discrepancies in the red slope lies in the different red/blue classification methods used in each work (and to a lesser extent, the differences in sample selection). If a different classification method results in a given percentage of 'contamination' compared to our classification (i.e. our blue galaxies being identified as red in another work, and vice-versa), this contamination will significantly affect the red sequence, as only 35 per cent of our sample is composed of red galaxies, according to our fiducial classification (i.e. see the histograms in Fig. 15). In comparison, the blue cloud will result mostly unaffected by contamination; hence the reasonable agreement between our D4000\(_v\)–mass blue slope and the results reported in Haines et al. (2017). The physical mechanism behind the increasing D4000\(_v\)–mass trend reported both here and in Haines et al. (2017), Siudek et al. (2017), is the downsizing scenario (Cowie et al. 1996), widely tested in the literature: more massive galaxies are older, and thus with higher D4000 values.

The D4000–SFR relation (see Fig. 15) is in good agreement with spectroscopy for both the PAUS direct measurement and the PAUS CIGALE reconstruction, with differences in median D4000 smaller than 0.1 for most cases (only a significant overestimation of \( \sim 0.2 \) appears in the lowest SFR bin). The MAD of the test data set (VIPERS D4000 with PAUS direct noise) is again similar to that of the PAUS direct measurement. The only significant caveat that arises is the underestimation of the MAD for the PAUS CIGALE D4000 reconstruction when compared to the VIPERS D4000 measurement, especially for the very low and highest SFR bins. This is consistent with our findings in the D4000–mass relation, and again shows that the SED templates do not properly reproduce the true galaxy diversity in the most quenched and star-forming regimes.

With the dependence of D4000 with galaxy properties shown in this subsection, and the similar trends found in the literature (with discrepancies stemming from different red/blue selection biases), we see that narrow-bands surveys show a great potential to break the incompleteness issue of flux-limited and target-based spectroscopy, as the direct photometric measurement of D4000 retrieves average trends compatible with spectroscopy. Moreover, the study of environmental dependence with statistical significance so far not achievable by spectroscopic surveys (Siudek et al. 2022) could also be performed with narrow-band data. The synergy of wide coverage with accurate measurements of spectral features may open a new insight in the nature versus nurture dilemma (e.g. Peng et al. 2010; De Lucia et al. 2012).

### 6 Conclusions

We have carried out the first in-depth study of the direct photometric measurement of the D4000 with narrow-band photometry, using the observational data of PAUS cross-matched with VIPERS (Scodellino et al. 2018) in the CFHTLS W1 field. The results of the photometric D4000 measurement have been compared to its spectroscopic counterpart, as well as the D4000 reconstructed from the SED-fitting code CIGALE (with either CFHTLS broad-bands or PAUS NBs). Both the original definition of the spectral break from Bruzual (1983): D4000\(_v\), and the narrower definition from Balogh et al. (1999): D4000\(_h\), have been considered.

First, we have developed a general D4000 estimator for narrow-band photometry, and we have tested it by measuring the D4000 in a synthetic PAUS catalogue, where we find that over 95 per cent (92 per cent) of objects have an absolute bias below 5 per cent in the PAUS D4000\(_v\), D4000\(_h\) measurement. We also find similar bias trends to the preliminary work of Stothert et al. (2018), but with a significant improvement (the average bias for D4000\(_h\) is reduced by ~2). With real PAUS narrow-band observations, only 85.01 per cent (D4000\(_v\)) and 65.87 per cent (D4000\(_h\)) have SNR > 3. However, an \( i_{AB} < 21 \) cut results in a bright sample where virtually all of the objects have SNR > 3. Hence, the PAUS catalogue achieves 'D4000-completeness' two magnitudes below the actual magnitude limit. The D4000 reconstructions from CIGALE present much higher SNR, but their errors are underestimated by >50 per cent. All estimation methods show similar bias trends: an absolute median bias of few per cent and a negative bias-VIPERS D4000 correlation. We also have compared D4000 distributions, and find that the PAUS direct D4000 follows closely the VIPERS D4000 with photometric Gaussian noise, and that the CIGALE reconstruction is far more realistic if PAUS NBs are used instead of CFHTLS broad-bands.

We have evaluated the D4000–redshift relation, and find that the PAUS direct measurements are fully compatible with VIPERS D4000, but the PAUS CIGALE D4000 overestimates the slope by a factor of ~2. Moreover, we have also examined the potential of these D4000 estimations for galaxy classification by comparing a given D4000\(_v\) value to a fiducial classification (Siudek et al. 2018). The D4000 reconstruction from PAUS CIGALE outperforms even
VIPERS spectroscopy (>90 per cent of correctly classified objects versus >85 per cent), while the PAUS direct measurements perform poorly due to the larger scatter (81 per cent of correctly classified objects at best). Finally, we have evaluated the D4000–stellar mass and D4000–SFR relation, where we find a generally good agreement between estimation methods. For the D4000–mass relation, we have separated the sample in red/blue galaxies, using the same fiducial classification, and find a blue slope reasonably close to previous results (Haines et al. 2017). The red slope seems to be steeper than in other works (e.g. Haines et al. 2017; Steidel et al. 2017), due to the different red/blue classifications employed.

In conclusion, this work shows that narrow-band photometry with PAUS-like spectral resolution ($R \sim 65$) can be used to directly measure the D4000 in a model-independent way; all objects two magnitudes below the magnitude limit can be individually measured with SNR > 3. For fainter magnitudes, stacking low-SNR measurements yields results consistent with spectroscopy. The D4000 reconstruction via SED fitting greatly benefits from the use of narrow-band photometry, providing a far more realistic distribution. However, it presents some artefacts such as noise underestimation, artificial bimodality, or underestimation of sample variance.

Therefore, these results open up several possibilities for narrow-band photometry that may not be available to conventional spectroscopy, such as the study of the radial dependence at different apertures (stacking several objects if necessary). More conventional studies can also be carried out, e.g. exploring the star formation history by constraining their stellar ages and redshift of formation. The wealth of narrow-band photometric data from PAUS may also be used to further study the relationship between morphology and D4000/galaxy quenching (e.g. see Kim et al. 2018, and the references therein), or to analyse the role of the environment in these processes, and thus disentangle the influence of nature versus nurture in galaxy evolution (e.g. Peng et al. 2010; De Lucia et al. 2012).

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DATA AVAILABILITY

The data underlying this article will be shared on reasonable request to the corresponding author.

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APPENDIX A: REDSHIFT COMPARISON

A comparison of the distribution of PAUS photo-z and VIPERS spec-z for the WI cross-match is shown in Fig. A1. The orange box in the middle of the plot represents the 0.562 < z < 0.967 cut in both redshift estimations; hence, all objects inside the box are the selected sample for D4000 measurements. Most of the PAUS photo-z follows closely along the red diagonal line representing the 1:1 equivalence. It is worth noting that the determination of PAUS photo-z in the wide fields is still a work in progress, and thus some artefacts that can be appreciated in Fig. A1 (horizontal stripes) may disappear in future data releases.

We can further evaluate the performance of PAUS photo-z with its dispersion, $\sigma_{\Delta z} = \sqrt{\langle (\Delta z)^2 \rangle}$, where $\Delta z$ is the difference between the two redshift determinations. For any given bin, we can calculate $\sigma_{\Delta z}$ as

$$
\sigma_{\Delta z} = \frac{\langle (\Delta z)^2 \rangle - \langle \Delta z \rangle^2}{1 + \langle \Delta z \rangle^2},
$$

where $\langle \Delta z \rangle$ is the average of $\Delta z$ for all objects in the bin, and $\langle (\Delta z)^2 \rangle$ is the average of the squares of $\Delta z$. Then, we have $\sigma_{\Delta z}^2 = \langle (\Delta z)^2 \rangle - \langle \Delta z \rangle^2$.

We can define $\sigma_{\Delta z}$ as

$$
\sigma_{\Delta z} = \frac{P_{\text{1.5} \text{ per cent}}(\Delta z) - P_{\text{95.9} \text{ per cent}}(\Delta z)}{2},
$$

where $P_{\text{1.5} \text{ per cent}}$ and $P_{\text{95.9} \text{ per cent}}$ are the 1.5th and 95.9th percentiles of the distribution, respectively.

Note that in practice, we usually use $\sigma_{\Delta z}$ as a measure of the uncertainty in the redshift, and $\sigma_{\Delta z}$ is a measure of the dispersion of the redshifts.

For the full redshift range, however, the outlier fraction of PAUS photo-z seems significantly larger than CFHTLS at $i_{\text{AB}} > 21.5$; this is mostly due to the $z < 1.2$ limit in the photo-z run, and does not affect the sample selected for this work. The $\sigma_{\Delta z}$ of PAUS photo-z also is slightly larger than its CFHTLS counterpart at the faint magnitude end ($i_{\text{AB}} > 22$); given that both PAUS NBs and CFHTLS broad-band data are used in the PAUS photo-z calculation (Eriksen et al. 2019), ideally the PAUS photo-z should never perform worse than CFHTLS alone. Hence, this higher PAUS $\sigma_{\Delta z}$ is an issue of the NB weighting relative to broad-band data, which is currently being worked on. Also, the worse performance of PAUS photo-z shown here when compared to earlier works in the COSMOS field ($\langle \sigma_{\Delta z}/(1 + z) \rangle \approx 0.0037$ for a 50 per cent quality cut; Eriksen et al. 2019; Alarcon et al. 2021) is mostly due to the smaller number of exposures and higher mean redshift in the wide fields.
Figure A1. PAUS photo-z versus VIPERS spec-z, for the cross-match of 28,788 galaxies with accurate spec-z measurements. Points are coloured according to the logarithm of the point density (with lighter areas having more objects). The red line represents the 1:1 equivalence between redshift measurements. The redshift cuts introduced for the D4000 measurement, both in PAUS photo-z and VIPERS spec-z, are represented by the orange box.

Figure A2. $\sigma_{68}/(1+z)$ (middle panel) and outlier fraction (lower panel) for the PAUS photo-z (orange lines) and CFHTLS photo-z (green lines), in magnitude bins. The solid line represents the full cross-match, dotted line the redshift cut for D4000 measurement. Upper panel displays the magnitude histogram.

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