A Population Pharmacokinetic and Pharmacodynamic Model of CKD-519

Choon Ok Kim 1, Sangil Jeon 2, Seunghoon Han 3,4, Min Soo Park 5,6 and Dong-Seok Yim 2,4,*

1 Department of Clinical Pharmacology, Severance Hospital, Yonsei University College of Medicine, Seoul 03722, Korea; delivery98@yuhs.ac (C.O.K.); minspark@yuhs.ac (M.S.P.)
2 Qfitter Inc., Seoul 06578, Korea; si.jeon@qfitter.com (S.J.); waystolove@catholic.ac.kr (S.H.)
3 Department of Pharmacology, Seoul St. Mary’s Hospital, College of Medicine, The Catholic University of Korea, Seoul 06591, Korea
4 PIPET (Pharmacometrics Institute for Practical Education and Training), College of Medicine, The Catholic University of Korea, Seoul 06591, Korea
5 Department of Pediatrics, Yonsei University College of Medicine, Seoul 03722, Korea
* Correspondence: yimds@catholic.ac.kr; Tel.: +82-2-2258-7357

Received: 5 June 2020; Accepted: 18 June 2020; Published: 19 June 2020

Abstract: CKD-519 is a selective and potent cholesteryl ester transfer protein (CETP) inhibitor that is being developed for dyslipidemia. Even though CKD-519 has shown potent CETP inhibition, the exposure of CKD-519 was highly varied, depending on food and dose. For highly variable exposure drugs, it is crucial to use modeling and simulation to plan proper dose selection. This study aimed to develop population pharmacokinetic (PK) and pharmacodynamics (PD) models of CKD-519 and to predict the proper dose of CKD-519 to achieve target levels for HDL-C and LDL-C using results from multiple dosing study of CKD-519 with a standard meal for two weeks in healthy subjects. The results showed that a 3-compartment with Erlang’s distribution, followed by the first-order absorption, adequately described CKD-519 PK, and the bioavailability, which decreased by dose and time was incorporated into the model (NONMEM version 7.3). After the PK model development, the CETP activity and cholesterol (HDL-C and LDL-C) levels were sequentially modeled using the turnover model, including the placebo effect. According to PK-PD simulation results, 200 to 400 mg of CKD-519 showing a 40% change in HDL-C and LDL-C from baselines was recommended for proof of concept studies in patients with dyslipidemia.

Keywords: population PK-PD modeling; CKD-519; CETP inhibitor; CETP activity; HDL-C

1. Introduction

Cholesteryl ester transfer protein (CETP) is a hydrophobic plasma glycoprotein consisting of 476 amino acids with a mass of 53 kDa [1]. It mediates bidirectional transfers of cholesteryl ester and triglyceride between high-density lipoprotein cholesterol (HDL-C) and apolipoprotein B-containing lipoproteins, such as very-low-density lipoprotein and low-density lipoprotein cholesterol (LDL-C) [2]. It has been reported that CETP plays a key role in atherosclerosis through its impact on lipid metabolism [1]. Thus, there have been many efforts to modify the composition of the plasma lipoprotein favorably through inhibition of CETP activity [3].

CKD-519 is a potent, selective CETP inhibitor being developed for the treatment of dyslipidemia. It is one of the cycloalkenyln benzene derivatives with a molecular weight of 601.60 g/mol (C31H34F7NO3) [4]. It is poorly soluble, with aqueous solubility of less than 1.0 μg/mL [4]. In the in vitro solubility test, CKD-519 was insoluble in deionized water and freely soluble in ethyl alcohol with its LogD of 7.4, and LogP is >5.0 at pH 7.4 [4,5].
CKD-519 exhibited a maximum of 63–83% inhibitory effect of CETP activity after single administration with 25–400 mg in healthy subjects (EC50 17.3 ng/mL) [5]. The exposure of CKD-519 increased with the dose. However, maximum plasma concentration (Cmax) and area under the plasma concentration-time curve (AUC) normalized by dose decreased with each incremental dose [5]. In addition, the pharmacokinetics (PK) of CKD-519 exhibited significant variation in the fasted and fed states. Administration of single CKD-519 200 mg with a standard meal resulted in an 8.7 and 5.8 fold increase in Cmax and AUC relative to the fasted state, respectively (in house data). The inhibitory effect of CETP activity by CKD-519 200 mg was also higher in the fed (91.4%) than in the fasted group (81.5%).

Despite potent CETP inhibition, the exposure of CKD-519 varied by food and dose. This study aimed to characterize the population PK and pharmacodynamics (PD) of CKD-519 using multiple dosing study of CKD-519 with a standard meal in healthy subjects. The PK-PD model can be applied to the design of proof of concept studies to observe HDL-C and LDL-C changes in patients.

2. Materials and Methods

2.1. Ethics

The protocols were approved by the institutional review board of Severance Hospital, Yonsei University College of Medicine (Seoul, Korea; IRB number 4-2015-0126 and 4-2016-0949, 02 APR 2015 and 17 DEC 2016). They were also registered at the clinicaltrials.gov (identifier: NCT02753504 and NCT03210649, 28 APR 2016 and 07 JUL 2017). These studies were carried out in accordance with the relevant regulatory requirements, including the Declaration of Helsinki and the Korean good clinical practice. All subjects provided written informed consent before enrollment in this study.

2.2. Study Design

A randomized, double-blinded, placebo-controlled, multiple-dose, and dose-escalation study was conducted in a total of 32 healthy male subjects aged 19–55 years in four cohorts (50, 100, 200, and 400 mg). Six subjects were assigned to CKD-519, and two to matching placebo in each cohort. Dose escalation to the next cohort was done after the safety profile was reviewed. CKD-519 was available as two dosage forms of 50 and 100 mg tablets.

Enrolled subjects were admitted on day −1. CKD-519 or placebo was administered to randomly assigned subjects for 14 days with a standard breakfast meal (700–800 kcal, containing 5–25% fat content). All subjects were hospitalized until completion of blood sampling for PK and PD (from day −1 to day 21). All meals were nutritionally equivalent and provided at the same time scheduled throughout the study period. The participants were not permitted to take any drugs or herbal medications and also were restricted from consuming any other snacks or beverages during the study.

Blood samples for PK or PD were collected on days 1 (0, 1, 2, 3, 4, 5, 6, 7, 8, 10, 12, 18, and 24 h) and 14 (0, 1, 2, 3, 4, 5, 6, 7, 8, 10, 12, 18, 24, 32, 48, 72, 96, 120, 144, and 168 h), along with daily sampling on days 2 through 13 to determine trough levels of drug concentration, CETP activity, and lipid values. Each blood sample for PK and CETP activity was collected in sodium heparin tubes and then centrifuged at 1900 g for 10 min. Their aliquots were stored at or below 70 °C until analysis. The plasma concentrations of CKD-519 were measured using a validated high-performance liquid chromatography-tandem mass spectrometry assay [5]. A 100 µL plasma sample was mixed with 10 µL internal standard working solution and 300 µL acetonitrile. After centrifugation, 3 µL supernatant was injected into the column. The lower limit of quantification was 1.0 ng/mL. The calibration curve was linear over the concentration range of 1.0–2000 ng/mL. The precision of the assay was less than 2.7% coefficient of variation, and the accuracy of the assay was within the range of 93.7–113.0% [5]. Plasma CETP activity was measured by using a fluorescent assay method previously described [5]. HDL-C and LDL-C were measured with an ADIVA 1650 Clinical Chemistry System (Siemens Medical Solutions, Tarrytown, NY, USA).
2.3. Model Development

PK and PD models were developed using NONMEM version 7.3 (Icon Development Solutions, Ellicott City, MD, USA) using a first-order conditional estimation with interaction method (FOCE-INTER). R (version 3.6.1, R Foundation for Statistical Computing, Vienna, Austria) was used for data preparation, graphical analysis, and model diagnostics. The population PK-PD model was developed sequentially. The PK model was initially developed, and then the CETP model (PK-PD) was developed using individual PK parameters estimated from the final PK model. Finally, the cholesterol (HDL-C or LDL-C) models were developed using the individual CETP parameters estimated from the final CETP model.

Inter-individual variability (IIV) for each parameter in all models was assumed to follow a Gaussian distribution with a mean of 0 and variance of \( \omega^2 \). For the PK model, once the IIV term was selected, the covariance between the terms was assessed using an omega block on selected parameters. As for the residual error, additive, proportional, and combined models were tested. The model was selected after a comparison of several criteria such as a decrease in the objective function value (OFV), completion of the estimation and covariance routines, the precision of the parameter and error estimates, and visual exploration of the diagnostic plots.

Once the structural model was evaluated, covariate analysis was performed using the stepwise forward selection and backward elimination process. Statistical significance was defined as a decrease in the OFV by more than 3.84 for forward selection and an increase in the OFV of at least 6.635 (chi-square distribution, 1 degree of freedom). The following factors were evaluated as covariates: age, weight, body mass index, creatinine clearance calculated from the modification of diet in renal disease equation, and alanine aminotransferase. They were considered significantly associated with the parameters if both of the OFV decrease and physiological plausibility were satisfied.

2.3.1. Population PK Model

Several disposition models were assessed, including 1-, 2-, and 3- compartment structural models. Absorption was evaluated with first-order, zero-order (with and without lag time), Weibull-type, and Erlang’s model [6]. The bioavailability was assessed about the effects of the dose and time. Drug elimination was assumed to follow first-order kinetics.

2.3.2. Population CETP Model

Since CETP inhibition is closely related to HDL-C increase and LDL-C decrease, the CETP model was included as a link between the PK model of CKD-519 and cholesterol (HDL-C and LDL-C) models [7]. For the CETP model, CETP activity was used as a dependent variable as this was the primary parameter of interest.

The turnover model was used to describe the relationship between CETP activity and plasma concentration of CKD-519. The effect of CKD-519 was applied in the model as stimulating the elimination of CETP activity. Because the CETP activity increased over time in the placebo group, the response of placebo treatment was incorporated into the CETP model. The effects of CKD-519 and placebo treatment were evaluated as linear, \( E_{max} \), or sigmoid \( E_{max} \) relationships.

2.3.3. Population HDL-C and LDL-C Models

For each of the HDL-C and LDL-C model development, their clinical laboratory results were used. Their responses by changing CETP activity were described using turnover models. Taking into account the physiological mechanism of CETP on lipid metabolism, the CETP effect was incorporated into the model as a stimulatory function of HDL-C elimination and an inhibitory function of LDL-C elimination. The effect of CETP activity on HDL-C or LDL-C was evaluated using linear, \( E_{max} \), or sigmoid \( E_{max} \) models.
2.4. Model Evaluation

A bootstrap resampling method was used to evaluate the stability and reliability of the final models. Bootstrap of 1000 resamplings was performed for each final model of PK, CETP, and HDL-C/LDL-C model to obtain the median and 95% confidence intervals for all parameters. In addition, the visual predictive check (VPC) with a total of 1000 simulated datasets was conducted on each PK, CETP, and HDL-C/LDL-C model. Results from the VPC were graphically assessed by comparing it with the overlaid observed data.

2.5. Simulation

The simulation was performed using the final model to guide the proper dose selection for an increase of HDL-C and decrease of LDL-C to the target levels. The target HDL-C and LDL-C levels were set by 40% changes compared to the baselines [8]. To simulate mean HDL-C and LDL-C values after three weeks of daily CKD-519 dosing, parameters of population fixed effects were drawn from final models. IVV or residual variability were not included in these simulations.

3. Results

3.1. Demographics

A total of 32 male subjects were randomized and completed these trials. Demographic and baseline characteristics are summarized in Table 1.

| Variables   | 50 mg   | 100 mg  | 200 mg  | 400 mg  | Placebo |
|-------------|---------|---------|---------|---------|---------|
|             | (n = 6) | (n = 6) | (n = 6) | (n = 6) | (n = 8) |
| Age (years) | 33.7 ± 5.5 | 34.0 ± 10.9 | 30.7 ± 9.9 | 30.3 ± 5.5 | 32.3 ± 9.7 | 0.924 |
| Height (cm) | 174.3 ± 7.8 | 174.4 ± 1.8 | 174.4 ± 6.7 | 175.1 ± 6.2 | 173.1 ± 4.9 | 0.980 |
| BMI (kg)    | 69.7 ± 7.8 | 69.8 ± 1.6 | 71.5 ± 8.2 | 67.7 ± 7.1 | 65.9 ± 7.6 | 0.609 |
| (kg/m²)     | 22.8 ± 1.0 | 22.9 ± 0.7 | 23.4 ± 1.8 | 22.0 ± 1.9 | 21.9 ± 2.2 | 0.436 |
| Cr (mg/dL)  | 0.79 ± 0.07 | 0.93 ± 0.05 | 0.89 ± 0.11 | 0.82 ± 0.09 | 0.84 ± 0.16 | 0.204 |
| ALT (IU/L)  | 14.0 ± 4.7 | 19.3 ± 6.9 | 13.5 ± 5.1 | 13.3 ± 11.1 | 14.4 ± 7.3 | 0.600 |

Data are given meansSD (min–max). BMI, body mass index; CR, creatinine; ALT, alanine aminotransferase. 'p-value among dose groups were calculated with the ANOVA test.

Demographic and baseline characteristics were well balanced across treatment groups. The mean age was 32.2 years (19.0 to 47.0 years), and the mean body weight was 68.7 kg (58.3 to 79.5 kg). A total of 1392 points of plasma CKD-519 concentration and 2064 points of CETP activity samples were measured. As for the HDL-C and LDL-C, 656 samples were measured.

3.2. Population Pharmacokinetic Analysis

The disposition of CKD-519 was assumed to follow first-order characteristics. A 3-compartment model with Erlang’s distribution (with five sequential compartments) followed by the first-order absorption best described the concentration data of CKD-519. A schematic diagram of the final PK model is shown in Figure 1.
Figure 1. Schematic representation of the final pharmacokinetic and pharmacodynamics models for CETP inhibitor. The dose (DOSE) goes into the absorption compartment via transit compartments. It described by Erlang’s absorption model with transit rate constant (Kα). The fraction of drug that reaches absorption compartment is dose and time-dependent. The formation of CETP activity described by the first-order CETP activity elimination rate (Kout), and also stimulated by the placebo effect via an Emax model. In addition, the elimination of CETP activity (Kout) stimulated by plasma drug concentration (Emax model). The response of HDL-C or LDL-C described by the zero-order production and first-order elimination. Also, the CETP effect was incorporated into the model as a stimulatory function of HDL-C elimination and an inhibitory function of LDL-C elimination (sigmoid Emax model).

Bmax, the maximal effect of dose on bioavailability; BA50, the dose of half-maximal effect on bioavailability; α, the proportionality constant for the fraction of first-order absorption process depending on time; Kν, transit rate constant; Kα, rate of constant for the first-order absorption; V/F, the volume of the central compartment; V2/F, the volume of peripheral compartment 1; V3/F, the volume of peripheral compartment 2; Q2/F, inter-compartmental clearance for compartment 1; Q3/F, inter-compartmental clearance for compartment 2; CL/F, clearance; Km, CETP activity production rate at baseline; Kmax, maximal effect of time of CETP activity production rate; K50, time of half-maximal effect on CETP production rate; Kout, the first-order rate constant for the elimination of CETP activity; Ksyn, the production rate of HDL-C or LDL-C; Kdeg, elimination rate of HDL-C or LDL-C; β, the proportionality constant for the fraction of LDL-C production process; T, time.

The following parameters were estimated in the model: absorption rate constant (Kα), transit rate constant (Kν), clearance (CL/F), central volume (V/F), peripheral compartment volumes (V2/F and V3/F), inter-compartmental clearances (Q2/F and Q3/F), and bioavailability (F).

A descriptive approach was used to describe the nonlinear exposure of CKD-519. The exposure of CKD-519 tended to decrease by increasing doses and repeated administration in the dose range of 50–200 mg. However, in the 400 mg group, its exposure after multiple administration resulted in a three-fold increase of AUC compared to that of the first dose. Therefore, the bioavailability terms were modeled using a combined term of dose and time as follows:

\[
F = BA \times FT 
\]  

(1)

\[
BA = B_{\text{max}} \left(1 - \frac{\text{DOSE}}{(BA_{50} + \text{DOSE})}\right) 
\]  

(2)
\[ FT = e^{-\alpha \times \text{TIME}} \]  

where \( BA \) is the relative bioavailability depending on dose, \( FT \) is the time-dependent absorption fraction, \( B_{\text{max}} \) is the maximal effect of dose on bioavailability, \( BA_{50} \) is the dose of a half-maximal effect on bioavailability, and \( \alpha \) is the proportionality constant for the fraction of absorption depending on time. The \( \alpha \) for 50–200 mg (\( \alpha_1 \)) was estimated in the model, but the \( \alpha \) for 400 mg (\( \alpha_2 \)) was fixed at zero.

Exponential IIV terms were included for \( \text{CL/F}, \text{V3/F}, B_{\text{max}}, \) and \( K_r \). Residual variability was represented by a proportional error. There was no statistically significant covariance included in the final model. Serum creatinine was identified as influential covariates for \( \text{CL/F} \) in the forward selection, but it did not satisfy backward elimination criteria. Covariate analysis was performed for the other variables, but none were significant. The final parameters estimated in the model are shown in Table 2.

**Table 2.** Parameter estimates from the final population pharmacokinetic model.

| Parameter | Description                                      | Estimate | %RSE | Bootstrap Median (95% CI)* |
|-----------|--------------------------------------------------|----------|------|---------------------------|
| \( CL/F \) (L/h) | Clearance                                      | 6.4      | 10.3 | 6.4 (5.2–8.1)             |
| \( V/F \) (L) | Volume of central compartment                      | 11.4     | 11.0 | 11.4 (9.4–15.1)          |
| \( V2/F \) (L) | Volume of peripheral compartment 1                  | 45.4     | 19.0 | 45.0 (23.2–72.2)         |
| \( V3/F \) (L) | Volume of peripheral compartment 2                  | 100.00   | 32.4 | 1060.0 (458.0–2320.3)    |
| \( Q2/F \) (L/h) | Inter-compartmental clearance for compartment 1 | 2.6      | 11.6 | 2.6 (1.9–3.6)            |
| \( Q3/F \) (L/h) | Inter-compartmental clearance for compartment 2     | 3.3      | 15.4 | 3.3 (2.6–4.8)            |
| \( K_r \) (h⁻¹) | Absorption rate constant                           | 1.09     | 3.5  | 1.09 (1.03–1.14)         |
| \( K_{\text{prop}} \) (h⁻¹) | Transit rate constant                             | 1.10     | 3.4  | 1.10 (1.03–1.16)         |
| \( F = BA \times FT \) |                                         |          |      |                           |
| \( B_{\text{max}} \) | Maximal effect of dose on bioavailability          | 1.6      | 30.7 | 1.7 (0.7–11.4)           |
| \( BA_{50} \) (mg) | Dose of half-maximal effect on bioavailability     | 90.1     | 24.5 | 87.8 (9.0–348.1)         |
| \( \alpha_1 \) | Proportionality constant for fraction of first-order absorption process depending on time for 50, 100, and 200 mg DOSE group | 0.002 | 14.4 | 0.002 (0.001–0.002) |
| \( \alpha_2 \) | Proportionality constant for the fraction of first-order absorption process depending on time for 400 mg DOSE group | 0 FIX | NA   | NA                       |

| Inter-Individual Variability (CV%) |
|------------------------------------|
| \( \omega_{\text{CL/F}} \) (%) | Interindividua variability on \( CL/F \) | 15.6 | 51.5 | 14.4 (0.3–29.2) |
| \( \omega_{\text{V3/F}} \) (%) | Interindividua variability on \( V3/F \) | 58.1 | 31.9 | 55.2 (20.0–77.2) |
| \( \omega_{\text{Bmax}} \) (%) | Interindividua variability on \( B_{\text{max}} \) | 28.2 | 41.6 | 26.1 (14.6–38.4) |
| \( \omega_{\text{Ktr}} \) (%) | Interindividua variability on \( K_r \) | 14.1 | 44.0 | 13.7 (6.4–19.5) |

| Residual Error |
|----------------|
| \( c_{\text{add}} \) | Additive error | 0.0001 FIX | NA | NA |
| \( g_{\text{group}} \) | Proportional error | 0.30 | 4.2 | 0.30 (0.27–0.32) |

*95% CI was estimated by applying the final population pharmacokinetic model to 1,000 resampled datasets. %RSE, relative standard error; CI, confidence interval; F, relative bioavailability; CV, coefficient of variation; NA, not applicable.
Figure 2A shows the model diagnostic plots for the final PK model. Residual errors in the lower panels are evenly scattered over the range of predicted concentrations. The VPC obtained with a 1000 simulated datasets is presented in Figure 3A. The observed data were fairly well predicted by the final model.

(A) (B) (C) (D)

Figure 2. Goodness-of-fit plots for the final population pharmacokinetic and pharmacodynamic models: (A) Plasma concentration of CKD-519; (B) CETP activity; (C) HDL-C; (D) LDL-C.
Black line is line of identity, and gray line is locally weighted regression smooth line. IWRES, individual weighted residuals.

\[
\frac{d \text{CETP}_{\text{activity}}}{dt} = K_{\text{in,base}} \times (1 + \text{Placebo}) - K_{\text{out}} \times (1 + \text{Drug})
\]
where $K_{in,base}$ is the zero-order constant for the increase of CETP activity at baseline, placebo is the placebo response to CETP activity, $K_{out}$ is the first-order rate constant for the decrease in CETP activity, and Drug is the effect of CKD-519 on CETP activity.

Several structural models for effects of drug and placebo were tested, including linear, $E_{max}$, and sigmoid $E_{max}$ models. Both of them were best described with $E_{max}$ model. The placebo and drug effects were as follows, respectively:

$$\text{Placebo} = \frac{K_{\text{max}} \times \text{time}}{K_{50} + \text{time}}$$  \hspace{1cm} (5)

$$\text{Drug} = \frac{E_{\text{max}} \times C_p}{EC_{50} + C_p}$$  \hspace{1cm} (6)

where $K_{\text{max}}$ is the maximal effect of time on CETP activity production rate, $K_{50}$ is the time of half-maximal effect on CETP production rate, $E_{\text{max}}$ is the maximal effect of plasma drug concentration on CETP activity, $EC_{50}$ is the plasma concentration of half-maximal effect on CETP activity, and $C_p$ is the plasma concentration of CKD-519.

The final model for CETP activity includes IIV on $K_{in,base}$ and the maximum drug effect, and the combined error structure was used for residual variability. There were no significant covariates in the model. The final parameter estimates are presented in Table 3, and the goodness-of-fit plots of the final model is shown in Figure 2B. Overall, the results suggested a reasonable fit of the model to observed CETP activity.

### Table 3. Parameter estimates for CETP model.

| Parameter           | Description                                      | Estimate | %RSE | Median (95% CI)* |
|---------------------|--------------------------------------------------|----------|------|-----------------|
| $\text{CETP}_{\text{base}}$ (pmole) | Baseline CETP activity                           | 350.0    | 1.9  | 346.0 (315.0–372.0) |
| $K_{in}$ (pmole/h)  | CETP activity increase rate constant at baseline | 164.0    | 0.9  | 149.0 (33.0–345.1) |
| $K_{\text{max}}$    | Maximal effect of time on CETP activity production rate | 9.6      | 2.2  | 9.8 (3.8–124.2)  |
| $K_{50}$ (h)        | Time of half-maximal effect on CETP production rate | 9700.0   | 17.3 | 8540.0          |
| $E_{\text{max}}$    | Maximal effect of plasma drug concentration on CETP activity | 18.2     | 0.8  | 20.0 (14.3–83.4) |
| $EC_{50}$ (ng/mL)   | Plasma drug concentration of half-maximal effect on CETP activity | 587.0   | 3.7  | 597.0 (441.0–2581.0) |
| IIV                 | Interindividual variability on $K_{in}$         | 19.7     | 0.9  | 17.4 (9.9–26.3)  |
| IIV                 | Interindividual variability on $E_{\text{max}}$ | 39.5     | 0.5  | 41.1 (25.8–61.1) |
| Residual Error      | Additive error                                  |          |      | 40.4 (26.1–47.4) |
|                     | Proportional error                              |          |      | 0.112 (0.058–0.237) |

*95% CI was estimated by applying the final population pharmacokinetic model to 1000 resampled datasets. %RSE, relative standard error; CI, confidence interval; $K_{in}$, production rate of CETP activity; Drug, drug effect on CETP activity; $C_p$, plasma drug concentration; CV, coefficient of variation; NA, not applicable.
3.4. Population HDL-C and LDL-C Analysis

HDL-C and LDL-C models were best described with the sigmoid $E_{\text{max}}$ function. The LDL-C in the placebo group tended to increase gradually over time. Thus, a linear model was used to describe the placebo effect on LDL-C after comparison with $E_{\text{max}}$ and sigmoid $E_{\text{max}}$ models. The changes in HDL-C and LDL-C concentration were described as follows:

$$\frac{d\text{HDL}}{dt} = K_{\text{syn}} - K_{\text{deg}} \times A_{\text{HDL}} \times (1 + \text{RES}_{\text{CETP}})$$

$$\frac{d\text{LDL}}{dt} = K_{\omega} \times (1 + \text{PLA}) - K_{\text{deg}} \times A_{\text{LDL}} \times (1 - \text{RES}_{\text{CETP}})$$

$$\text{RES}_{\text{CETP}} = \frac{\text{CETP}\gamma_{\text{activity}} \times R_{\text{max}}}{\text{CETP}\gamma_{\text{activity}} + R_{50}}$$

$$\text{PLA} = \beta \times t$$

where $K_{\text{syn}}$ and $K_{\omega}$ are the zero-order production rate constant, $K_{\text{deg}}$ is the first-order elimination rate constant, $\text{RES}_{\text{CETP}}$ is the change in CETP activity, $R_{\text{max}}$ is the maximal effect of CETP activity on HDL-C and LDL-C response, $R_{50}$ is the CETP activity required to attain 50% maximal response of HDL-C and LDL-C, PLA is the placebo effect, and $\beta$ is the proportionality constant.

The final model of HDL-C included IIV on baseline HDL-C, and its residual error was additive. For the LDL-C model, IIV was used in the baseline LDL-C, $\beta$, and $\gamma$. The residual variability of LDL-C was in the combined structure. No covariates were found to be significant in the HDL-C and LDL-C models. The parameters estimated in the final model are shown in Table 4.

| Parameter | Description | Estimate | %RSE  | Bootstrap Median (95% CI)* |
|-----------|-------------|----------|-------|--------------------------|
| **HDL-C Structural Model** | | | | |
| RB (mg/dL) | Baseline HDL-C | 50.0 | 13.2 | 50.0 (46.3–53.5) |
| $K_{\text{syn}}$ (h$^{-1}$) | Production rate of HDL-C | 1.26 | 15.2 | 1.26 (1.03–1.57) |
| $\gamma$ | Steepness of the CETP activity versus HDL-C relationship | 2.1 | 5.3 | 2.1 (1.1–3.6) |
| $R_{\text{max}}$ | Maximal effect of CETP activity on HDL-C response | 1.5 | 56.5 | 1.6 (1.1–3.4) |
| $R_{50}$ (pmole) | CETP activity of half-maximal effect on HDL-C response | 185.0 | 24.8 | 185.0 (143.0–390.1) |
| $\omega_{R\text{B}}$ (%) | Inter-individual variability on RB Residual Error | 17.2 | 29.3 | 16.8 (12.2–21.3) |
| $\sigma_{\text{add}}$ | Additive error | 4.6 | 6.2 | 4.5 (4.0–5.0) |
| $\sigma_{\text{prop}}$ | Proportional error | 0.00001 | NA | NA |
| **LDL-C Structural Model** | | | | |
| RB (mg/dL) | Baseline LDL-C | 97.4 | 1.9 | 94.3 (83.4–105.0) |
| $K_{\text{SYN}}$ | $K_{\text{syn}} \times (1 + \beta \times T)$ | | | |
| $K_{\text{syn}}$ (h$^{-1}$) | Production rate of LDL-C | 0.435 | 1.5 | 0.419 (0.245–0.573) |
| $\beta$ | Proportionality constant for fraction of LDL-C production rate | 0.0009 | 0.8 | 0.0008 (0.0004–0.0021) |
| $\text{RES} = (\text{CETP}\gamma_{\text{activity}} \times R_{\text{max}})/(\text{CETP}\gamma_{\text{activity}} + R_{50})$ | | | | |
showed for responses respectively. 3.5.

As Figure 2C,D show the model diagnostic plots for our final HDL-C and LDL-C model, respectively. VPC from the 1,000 simulated datasets are presented in Figure 3C,D. The VPC results showed that the 90% prediction intervals of the simulated data matched well with the observed data for both of the HDL-C and LDL-C.

3.5. Simulation

The final HDL-C and LDL-C model were used in population simulations to find out the proper dose to achieve target lipid levels. The simulations predicted HDL-C and LDL-C after 21 days of administration of CKD-519 with different doses (Figure 4).

(A) HDL-C response (B) LDL-C response. They are simulated HDL-C and LDL-C responses after 21-day repeated administration of CKD-519. The dashed lines are the target values, which are set by a 40% change from baseline of HDL-C and LDL-C, respectively.

As the dose increased, the HDL-C and LDL-C responses also increased. They showed maximum responses after repeated dosing for ten days. Because the final model explained observed data using a time-dependent decrease of bioavailability for 50–200 mg and time-dependent increase of CETP activity (placebo group), the simulated effects of CKD-519 on cholesterol decreased gradually
afterward in 50–200 mg dose groups, but not in the 400 mg. The HDL-C increase of 40% and LDL-C decrease of 40% from baselines were observed in 200 mg and 400 mg groups.

4. Discussion

CKD-519 has shown variable systemic exposure depending on dose and food (in house data) [5]. It is necessary to understand the relationship between PK of CKD-519 and lipid response in order to find out an appropriate dose for the treatment of dyslipidemia. The PK-PD models of CKD-519 and HDL-C/LDL-C levels mediated by CETP activity were developed in this study in that context. At first, a PK/PD model was constructed to connect the CKD-519 exposure and CETP activity, and then, lipid responses (HDL-C and LDL-C levels) were further connected to the PK-CETP activity model.

In healthy subjects, CKD-519 exposure was found to increase in a less than dose-proportional manner with respect to dose over the 25–400 mg dose range that was studied [5]. In our report, less than dose-proportional increases in the CKD-519 exposure were also observed, with CL/F approximately 2.5-fold increase from 11.98 L/h at 50 mg to 30.04 L/h at 200 mg. However, CL/F for 400 mg was 23.98 L/h that was less than that of 200 mg. The observed terminal half-life did not appear to vary with dose, so the lack of dose-proportionality is thought to be the result of changes in the extent of absorption rather than changes in the rate of elimination as reported in other CETP inhibitors, anacetrapib, and evacetrapib [9]. The CKD-519 exposure decreased after multiple administration in the 50–200 mg dose range, but not in 400 mg. The saturation of the absorption pathway suggested in previous studies [5,10], or autoinduction [11] may be employed to account for the phenomena observed in 50–200 mg groups, but not in 400 mg. Thus, a descriptive model was applied to estimate relative bioavailability changes of CKD-519 with dose and time.

CKD-519 showed a tri-phasic disposition pattern and had a long half-life with 145.4–166.2 h. This feature was also similar in other CETP inhibitors [12]. Among them, anacetrapib was well known for its long terminal half-life. [13] Some studies reported that anacetrapib accumulates in the adipose tissue, and its concentration in blood declines very slowly [13,14]. Based on these findings, anacetrapib was also analyzed using a 3-compartment model [13]. Similar to anacetrapib, CKD-519 has neutral lipophilic properties, which explains its high affinity and accumulation in the adipose tissue [5]. Therefore, the PK model of CKD-519 employed a 3-compartment structure like in anacetrapib, which depicted the disposition of observed data better than 1- or 2-compartment structures.

Also, all dose groups of CKD-519 in our study have shown delayed absorption (Tmax approximately 1 h). Hence, a few delayed absorption models were tested under the 3-compartment structure. The Weibull-type absorption model used in the animal model of CKD-519 was unstable and terminated without minimization [10]. When Erlang’s absorption model was finally applied, OFV was the lowest, and goodness-of-fit plots were the best compared to other delayed absorption models. Five transit compartments were added, and the OFV was the lowest at 8789.41, which was reduced by about 386 compared to the OFV of zero-order with lag time model. According to these results, a 3-compartment PK model of CKD-519 with Erlang’s absorption (with five sequential compartments) was finally chosen.

CETP activity was rapidly suppressed and sustained during repeated CKD-519 administration. On the other hand, CETP activity gradually increased over time in the placebo group. As reported earlier in other CETP inhibitors [15], the administration of CKD-519 in this study led to an approximately 2.5 fold increase in serum CETP concentration in a time-dependent manner (in house data). In the placebo group, serum CETP concentration increased by about 30% after two weeks (in house data). It has been reported that transcription of the CETP gene is under the control of extrinsic and intrinsic factors [16]. Dietary cholesterol up-regulates CETP expression in human CETP- transgenic mice [16], and plasma cholesterol levels also correlate with CETP mass in humans [17]. The increase in CETP activity in the placebo group might be due to increased serum CETP mass because dietary cholesterol regularly served during hospitalization for three weeks could have promoted CETP expression. Besides, total cholesterol, including LDL-C and triglyceride in the
placebo group, also slowly increased, but their HDL-C levels were stable. Because increased plasma cholesterol could have up-regulated the CETP gene expression in the placebo group, the placebo effect term on CETP activity was tested in the PD modeling step. Linear, Emax, and sigmoid Emax models were tested to explain CETP activity increase in the placebo group, and the Emax model by time best described the placebo effect.

HDL-C and LDL-C responses to CKD-519 were rather slow, whereas the CETP activity rapidly decreased and stayed at the nadir. Turnover models were used to describe the relationship between CETP activity and HDL-C/LDL-C responses. In addition, a placebo effect model aforementioned was also used for LDL-C. The maximal increase in HDL-C was estimated to be 160–190% and serum CETP activity exerting the half-maximal effect (R50) on HDL-C was 185.0 pmole. The maximal decrease in LDL-C was estimated to be 47–66%, and the corresponding R50 was 80.7 pmole.

Using the final PK-CETP-lipid model, we performed simulations to investigate the proper dose of CKD-519 to achieve target levels for HDL-C (40% increase) and LDL-C (40% decrease) as in the case of anacetrapib [8]. The simulation showed that 200 to 400 mg CKD-519 once daily will result in lipid-altering effects that are near the PD targets (dotted lines, Figure 4). In the case of 50–200 mg, the simulated responses decreased slowly after the tenth day of dosing because of the time dependency of bioavailability and CETP activity. As our PK and PD data were obtained from rather a small number of young, healthy subjects admitted for only three weeks for a phase I clinical trial, such a time-dependency may not be generalizable to the clinical settings in patients. Moreover, patients with dyslipidemia have a 30% higher baseline level of LDL-C than healthy subjects, which implies the difference in lipid metabolism between healthy subjects and patients [8].

5. Conclusions

In conclusion, the PK and PK/PD relationships of CKD-519 have been well characterized using the data from a multiple dosing study in healthy subjects. As a result of simulation from the final CKD-519 model, dosage regimens of 200–400 mg daily were recommended for further proof of concept studies in patients.

Author Contributions: conceptualization: C.O.K., S.J., S.H., M.S.P., and D.S.Y. Conduct of clinical study: C.O.K. and M.S.P. Data curation: C.O.K. and S.J. Formal analysis: C.O.K., S.J., S.H., and D.S.Y. writing—original manuscript: C.O.K. and D.S.Y. All authors have read and agreed to the published version of the manuscript.

Funding: This research received no external funding.

Acknowledgments: This study was supported by Chong Kun Dang Pharmaceutical Corp., Seoul, Korea.

Conflicts of Interest: The authors declare no conflict of interest. The co-authors Sangil Jeon, Seunghoon Han, Min Soo Park, and Dong-seok Yim are employees of Q-fitter. The company had no role in the design of the study; in the collection, analyses, or interpretation of data; in the writing of the manuscript, or in the decision to publish the results.

References
1. Chowniecz, Z.; Skoczyńska, A. Plasma lipid transfer proteins: The role of pltp and ctep in atherogenesis. Adv. Clin. Exp. Med. 2018, 27, 429–436.
2. Barter, P.J.; Hopkins, G.J.; Calvert, G.D. Transfers and exchanges of esterified cholesterol between plasma lipoproteins. Biochem. J. 1982, 208, 1–7.
3. Grabie, M.; Tai, C.; Frishman, W.H. Is anacetrapib better than its ctep inhibitor counterparts? Cardiol. Rev. 2019, 27, 242–248.
4. Park, S.J.; Thapa, P.; Seo, H.; Park, E.S.; Jeong, S.H. Early formulation development of ckd-519, a new ctep inhibitor, for phase 1 clinical study based on in vitro and in vivo evaluation. Int. J. Pharm. 2018, 549, 388–396.
5. Kim, C.O.; Oh, E.S.; Choi, C.; Kim, Y.; Lee, S.; Kim, S.; Park, M.S. Pharmacokinetics, pharmacodynamics and safety of ckd-519, a ctep inhibitor, in healthy subjects. Drug. Dev. Dev. Ther. 2016, 10, 3763.
6. Rousseau, A.; Léger, F.; Le Meur, Y.; Saint Marcoux, F.; Paintaud, G.; Buchler, M.; Marquet, P. Population pharmacokinetic modeling of oral cyclosporin using nonmem: Comparison of absorption pharmacokinetic
models and design of a bayesian estimator. Ther. Drug. Monit. 2004, 26, 23–30.
7. Klerkx, A.H.; El Harchaoui, K.; van der Steeg, W.A.; Boekholdt, S.M.; Stroes, E.S.; Kastelein, J.J.; Kuivenhoven, J.A. Cholesteryl ester transfer protein (ctep) inhibition beyond raising high-density lipoprotein cholesterol levels: Pathways by which modulation of ctep activity may alter atherogenesis. Arterioscler. Thromb. Vasc. Biol. 2006, 26, 706–715.
8. Krishna, R.; Bergman, A.J.; Green, M.; Dockendorf, M.F.; Wagner, J.A.; Dykstra, K. Model-based development of anacetrapib, a novel cholesteryl ester transfer protein inhibitor. AAPS J. 2011, 13, 179–190.
9. Friedrich, S.; Kastelein, J.J.; James, D.; Waterhouse, T.; Nissen, S.E.; Nicholls, S.; Kuivenhoven, J.A. The pharmacokinetics and pharmacokinetic/pharmacodynamic relationships of evacetrapib administered as monotherapy or in combination with statins. CPT: PSP. 2014, 3, e94.
10. Choi, S.; Han, S.; Jeon, S.; Yim, D. Quantitative prediction of human pharmacokinetics and pharmacodynamics of ckd519, a potent inhibitor of cholesteryl ester transfer protein (ctep). Pharmaceutics 2019, 11.
11. Svensson, R.J.; Aarnoutse, R.E.; Diacon, A.H.; Dawson, R.; Gillespie, S.H.; Boeree, M.J.; Simonsson, U.S.H. A population pharmacokinetic model incorporating saturable pharmacokinetics and autoinduction for high rifampicin doses. Clin. Pharmacol. Ther. 2018, 103, 674–683.
12. Ferri, N.; Corsini, A.; Sirtori, C.R.; Ruscica, M. Present therapeutic role of cholesteryl ester transfer protein inhibitors. Pharmaco. Res. 2018, 128, 29–41.
13. Small, D.S.; Ke, A.B.; Hall, S.D.; Mantlo, N.; Rotelli, M.; Friedrich, S. Assessment of the persistence of anacetrapib and evacetrapib concentrations using two pharmacokinetic modeling approaches. J. Clin. Pharmacol. 2015, 55, 757–767.
14. Krishna, R.; Ghayas, F.; Liu, Y.; Hagen, D.R.; Walker, B.; Chawla, A.; Cote, J.; Blaustein, R.O.; Gutstein, D.E. Chronic administration of anacetrapib is associated with accumulation in adipose and slow elimination. Clin. Pharmacol. Ther. 2017, 102, 832–840.
15. Krishna, R.; Bergman, A.J.; Jin, B.; Fallon, M.; Cote, J.; Van Hoydonck, P.; Laethem, T.; Gendrano, I.N.; Van Dyck, K.; Hilliard, D. et al. Multiple-dose pharmacodynamics and pharmacokinetics of anacetrapib, a potent cholesteryl ester transfer protein (ctep) inhibitor, in healthy subjects. Clin. Pharmacol. Ther. 2008, 84, 679–683.
16. Shrestha, S.; Wu, B.J.; Guiney, L.; Barter, P.J.; Rye, K. Cholesteryl ester transfer protein and its inhibitors. J. Lipid Res. 2018, 59, 772–783.
17. Martin, L.J.; Connelly, P.W.; Nancoo, D.; Wood, N.; Zhang, Z.J.; Maguire, G.; Quinet, E.; Tall, A.R.; Marcel, Y.L.; McPherson, R. Cholesteryl ester transfer protein and high density lipoprotein responses to cholesterol feeding in men: Relationship to apolipoprotein e genotype. J. Lipid Res. 1993, 34, 437–446.

© 2020 by the authors. Licensee MDPI, Basel, Switzerland. This article is an open access article distributed under the terms and conditions of the Creative Commons Attribution (CC BY) license (http://creativecommons.org/licenses/by/4.0/).