Dipeptidyl peptidase-4 inhibitor cardiovascular safety in patients with type 2 diabetes, with cardiovascular and renal disease: a retrospective cohort study

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Clinical trials investigating cardiovascular safety of dipeptidyl peptidase-IV inhibitors (DPP-4i) among patients with cardiovascular and renal disease rarely recruit patients with renal impairment, despite associations with increased risk for major adverse cardiovascular events (MACE). We investigated the risk of MACE associated with the use of DPP-4i among these high-risk patients. Using a new-user, retrospective, cohort design, we analyzed 2010–2015 IBM MarketScan Commercial Claims and Encounters for patients with diabetes, comorbid with cardiovascular disease and/or renal impairment. We compared time to first MACE for DPP-4i versus sulfonylurea and versus metformin. Of 113,296 individuals, 9146 (8.07%) were new DPP-4i users, 17,481 (15.43%) were new sulfonylurea users, and 88,596 (78.20%) were new metformin users. Exposure groups were not mutually exclusive. DPP-4i was associated with lower risk for MACE than sulfonylurea (aHR 0.84; 95% CI 0.74, 0.93) and similar risk for MACE to metformin (aHR 1.07; 95% CI [1.04, 1.16]). DPP-4i use was associated with lower risk for MACE compared to sulfonylureas and similar risk for MACE compared to metformin. This association was most evident in the first year of therapy, suggesting that DPP-4i is a safer choice than sulfonylurea for diabetes treatment initiation in high-risk patients.

Diabetes afflicts nearly 26 million people in the United States and is associated with considerable morbidity, mortality, and health care spending1,2. This statistic becomes especially concerning given that many patients with diabetes later develop diabetic complications such as diabetic retinopathy, diabetic ketoacidosis, and peripheral vascular disease. Additionally, patients with diabetes also often suffer from cardiovascular disease3. As such, patients and providers are left with a host of considerations for managing both the symptoms of diabetes and the associated comorbidities.

One important class of medicines to treat type 2 diabetes are the dipeptidyl peptidase-IV inhibitors (DPP-4i), which act by slowing the breakdown of glucagon-like peptide-1 (GLP-1), inhibiting glucagon release and increasing insulin release4. Since they were introduced in 2006, four DPP-4i have been approved by the FDA for stand-alone use or as part of fixed-dose combination products, and they rank third in utilization after metformin and sulfonylurea with 8% of antidiabetic drug prescriptions5.

DPP-4i were initially believed to be protective against major adverse cardiovascular events (MACE), evidenced through pre-market clinical trials6. Given high rates of cardiovascular disease among patients with...
diabetes\textsuperscript{7,8}, as well as longstanding regulatory interest in the potential adverse cardiovascular events associated with diabetes treatments\textsuperscript{9}, evidence of such a cardioprotective effect was of high regulatory, clinical and market importance. However, despite this early evidence from pre-market clinical pharmacology studies, postmarketing surveillance reports\textsuperscript{10,11} and data from Phase 4 clinical trials\textsuperscript{12–14} suggested a possible harmful association with MACE, specifically heart failure.

We examined the association between DPP-4i therapy and cardiovascular events using a large commercial claims database. We focused on individuals at elevated baseline risk, including those with history of cardiovascular disease as well as those with renal impairment, comparing the rates of adverse cardiovascular events among DPP-4i new users with that of metformin and sulfonylurea. The inclusion of individuals with renal impairment reflected increased regulatory interest in understanding cardiovascular risk in this patient population as evidenced in the two latest clinical trials studying the association of DPP-4i and cardiovascular outcomes\textsuperscript{15,16}. Additionally, the results of the Saxagliptin Assessment of Vascular Outcomes Recorded in Patients with Diabetes Mellitus Thrombolysis in Myocardial Infarction 53 (SAVOR-TIMI 53) trial showed that an elevated urinary albumin to creatinine ratio was independently associated with increased risk for MACE\textsuperscript{17}. We were interested in this elevated risk population, because if a discernable effect were present, it could mean additional cardiovascular risk from DPP-4i therapy for a population already at high susceptibility to cardiovascular events.

**Results**

We identified 12,166,812 individuals with diabetes from 2010 to 2015. After applying our inclusion and exclusion criteria, there were a total of 113,296 individuals in our cohort (Fig. 1). Of these, 9146 (8.07%) were new users of DPP-4i, 17,481 (15.43%) initiated sulfonylureas, and 88,596 (78.20%) started metformin. Three percent of included individuals were new users of both DPP-4i and metformin, and 1524 (1.72%) were new users of both sulfonylureas and metformin.

Table 1 depict the demographic and clinical characteristics of DPP-4i users and their counterparts. For example, more sulfonylurea users (60.40%) were male compared to DPP-4i (58.51%) or metformin (52.17%) users.

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1 OHA = oral antihyperglycemic agents
2 Subcategories are not mutually exclusive

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**Figure 1.** Cohort derivation and sample attrition after applying inclusion and exclusion criteria for retrospective cohort.
Approximately half of individuals in each exposure group had a cumulative exposure to the treatment of interest less than or equal to 6 months (DPP-4i: 54.05%; sulfonylureas: 56.28%; metformin: 53.65%). There were also several differences in baseline comorbidities, and corresponding medication use, between exposure groups. For example, rates of cardiovascular disease were higher among users of metformin than their counterparts, while

Table 1. Baseline demographics and medical characteristics. *DCSI adapted diabetes complications severity index. Includes diseases of kidneys, ureters, and bladders; inclusive of acute renal failure.

| Baseline Characteristics | DPP-4I (n = 9146) | Sulfonylureas (n = 17,481) | Metformin (n = 88,596) |
|--------------------------|------------------|--------------------------|------------------------|
| Male sex                 | 5351 58.51       | 10,559 60.40             | 46,224 52.17           |
| Age, years               |                  |                          |                        |
| 35–44                    | 877 9.59         | 2066 11.82               | 11,755 13.27           |
| 45–54                    | 2815 30.78       | 5147 29.44               | 29,238 33.00           |
| 55–64                    | 5454 59.63       | 10,268 58.74             | 47,603 53.73           |
| Cumulative exposure, months |               |                          |                        |
| < 6 months               | 4943 54.05       | 9838 56.28               | 47,534 53.65           |
| 6–12 months              | 2580 28.21       | 4326 24.75               | 23,869 26.94           |
| 12–18 months             | 848 9.27         | 1665 9.52                | 8713 9.83              |
| > 18 months              | 775 8.47         | 1652 9.45                | 8480 9.57              |
| Comorbidities in baseline |               |                          |                        |
| Cardiovascular disease   | 8187 89.51       | 15,359 87.86             | 84,138 94.97           |
| Kidney disease (chronic and acute) | 1372 15.00    | 3011 17.22               | 6263 7.07              |
| Cerebrovascular disease  | 1574 17.21       | 2931 16.77               | 14,391 16.24           |
| Ischemic heart disease   | 3926 42.93       | 7374 42.18               | 35,242 39.78           |
| Hypertension             | 7501 82.01       | 13,902 79.53             | 68,723 77.57           |
| Eye disease              | 3149 34.43       | 5289 30.26               | 28,809 32.52           |
| Renal disease*           | 4029 44.05       | 7243 41.43               | 34,718 39.19           |
| Acute renal failure      | 465 5.08         | 1126 6.44                | 2118 2.39              |
| Neuropathy               | 1320 14.43       | 2270 12.99               | 12,573 14.19           |
| Nephropathy              | 559 6.11         | 1330 7.61                | 2455 2.77              |
| aDCSI score              |                  |                          |                        |
| 0                        | 5813 63.56       | 11,326 64.79             | 61,295 69.18           |
| 1                        | 1203 13.15       | 2055 11.76               | 10,806 12.20           |
| 2                        | 1358 14.85       | 2738 15.66               | 11,290 12.74           |
| 3 +                      | 772 8.44         | 1362 7.79                | 5205 5.87              |
| Dual exposures           |                  |                          |                        |
| Sulfonylureas            | 0 0.00           | 0 0.00                   | 1,524 1.72             |
| Metformin                | 3451 37.73       | 1524 8.72                | 0 0.00                 |
| DPP-4 inhibitors         | 0 0.00           | 0 0.00                   | 3451 3.90              |
| Concomitant baseline medications |              |                          |                        |
| ACE inhibitors           | 2022 22.11       | 3643 20.84               | 20,522 23.16           |
| Angiotensin II receptor blockers | 2380 26.02    | 3346 19.14               | 18,410 20.78           |
| Antiadpressants          | 2132 23.31       | 3378 19.32               | 24,832 28.03           |
| Antiplatelets            | 2871 31.39       | 4729 27.05               | 28,182 31.81           |
| Asthma medication        | 4599 50.28       | 7531 43.08               | 48,709 54.98           |
| Benzodiazepines          | 1842 20.14       | 2905 16.62               | 20,485 23.12           |
| Beta blockers            | 2832 30.96       | 5208 29.79               | 28,437 32.10           |
| Blood thinners and anticoagulants | 496 5.42    | 820 4.69                 | 4924 5.56              |
| Calcium channel blockers | 1507 16.48       | 2762 15.80               | 13,754 15.52           |
| Cardioselective beta blockers | 940 10.28    | 1632 9.34                | 7767 8.77              |
| Diuretics                | 1737 18.99       | 3016 17.25               | 19,258 21.74           |
| Nitrites                 | 807 8.82         | 1342 7.68                | 7924 8.94              |
| Peripheral neuropathic treatments | 688 7.52     | 944 5.40                 | 6867 7.75              |
| Statins                  | 4699 51.38       | 7484 42.81               | 41,933 47.33           |
| Thiazide diuretics       | 1875 20.50       | 3335 19.08               | 21,306 24.05           |
kidney disease was more prevalent among users of sulfonylureas (17.22%) and DPP-4i (7.07%). Differences between users of DPP-4i, sulfonylurea, and metformin were not statistically significant after application of propensity score weighting.

The median follow-up for the primary outcome was 160 days (interquartile range (IQR) 92–296 days) and there was no statistically significant difference in time-to-event distributions across the therapeutic classes examined. After propensity score weighting and adjusting for baseline demographics, clinical characteristics, and concomitant medications, the incidence rate for the primary outcome was statistically significantly less (30.20 per 1000 person-years) among new users of DPP-4i compared to sulfonylureas (Table 2). There was a non-statistically significant increase in the incidence rate for the primary outcome among new users of DPP-4i compared to sulfonylureas (91.25 per 1000 person-years vs. 79.46 per 1000 person-years). Among the secondary outcomes, the incidence rate for heart failure was statistically significantly less (4.07 per 1000 person-years vs. 7.56 per 1000 person-years). There were no statistically significant differences between the overall and cumulative exposure stratum specific adjusted hazard ratios for MACE among users of DPP-4i compared to sulfonylurea or metformin (Additional File 1, Supplementary Appendix 3).

When comparing the association between DPP-4i and the primary outcome by cumulative exposure strata, there were differences from the overall effect. In the comparison between new users of DPP-4i and sulfonylurea, the adjusted hazard ratio for the primary outcome was statistically significant for a cumulative exposure of < 6 months (aHR 0.85, 95% CI 0.71–0.93) and 6–12 months (aHR 0.76, 95% CI 0.70–0.93). However, statistical significance was not evident for cumulative exposure 12–18 months and > 18 months. This suggests that the overall aHR might be driven by differences in MACE risk seen in the first year after initial exposure. There were no differences between the overall and cumulative exposure stratum specific adjusted hazard ratios for MACE between new users of DPP-4i and metformin (Table 4).

| Outcome, n | DPP-4 inhibitors | Sulfonylureas | Metformin |
|------------|-----------------|-------------|-----------|
| Primary composite outcome | 510 | 1301 | 4781 |
| Total person years | 5578 | 10,696 | 57,134 |
| Rate per 1000 person years | 91.25 | 121.63 | 79.46 |
| Median [IQR] observation time, days | 160 [92, 286] | 144 [75, 288] | 163 [93, 300] |
| Rate difference (95% CI) | − | − | 30.20 [− 40.47, − 19.89] |
| Acute myocardial infarction | 54 | 132 | 450 |
| Total person years | 5898 | 11,487 | 60,159 |
| Rate per 1000 person years | 9.16 | 11.49 | 7.06 |
| Median [IQR] observation time, days | 167 [98, 297] | 161 [90, 308] | 172 [100, 314] |
| Rate difference (95% CI) | − | − | 2.34 [− 5.45, 0.80] |
| Stroke | 141 | 268 | 1202 |
| Total person years | 5839 | 11,380 | 59,592 |
| Rate per 1000 person years | 24.15 | 23.55 | 20.17 |
| Median [IQR] observation time, days | 167 [97, 295] | 159 [90, 305] | 171 [99, 311] |
| Rate difference (95% CI) | − | − | 0.60 [− 4.31, 5.47] |
| Heart failure | 24 | 87 | 240 |
| Total person years | 5912 | 11,512 | 60,277 |
| Rate per 1000 person years | 4.06 | 7.56 | 3.98 |
| Median [IQR] observation time, days | 168 [98, 298] | 161 [90,308] | 172 [100, 314] |
| Rate difference (95% CI) | − | − | 3.50 [− 5.78, − 1.21] |

Table 2. Incidence rates of primary composite outcome, acute myocardial infarction, stroke, and heart failure among new users of DPP-4 inhibitors, sulfonylureas, and metformin. IQR interquartile range.

1Primary composite outcome includes myocardial infarction, cardiac arrest, coronary artery bypass, coronary angioplasty, heart failure, stroke, inpatient death. 2Incidence rate difference per 1000 person-years. Bold values indicate statistically significant results.
Results of analyses excluding individuals with acute renal failure (sulfonylureas: aHR 0.83, 95% CI 0.17–0.86; metformin: aHR 1.08, 95% CI 1.01–1.19) showed qualitatively similar results to those including these individuals (sulfonylureas: aHR 0.84, 95% CI 0.74–0.93; metformin: aHR 1.07, 95% CI 0.98–1.16). Similarly, results were not sensitive to changes in the latency period after the last dose of exposure for the primary analysis results, nor did they differ substantively when lagging the latency period by 7-days or 30-days (Additional File 1, Supplementary Appendix 4).

Discussion
In this longitudinal study of commercial claims data for individuals with diabetes, having also cardiovascular disease and renal impairment, new use of DPP-4i was associated with a lower risk for MACE compared to sulfonylureas, and a comparable risk compared to metformin. This difference was most evident in the first year of use. New use of DPP-4i was not shown to be associated with the following individual components of MACE: heart failure, stroke, or acute myocardial infarction. These results contribute to the body of evidence examining the association of DPP-4i and MACE in individuals at higher risk for cardiovascular events.

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Table 3. Hazard ratios for the association between DPP-4 inhibitor use and primary composite outcome, acute myocardial infarction, stroke, and heart failure compared to sulfonylureas and metformin. *Propensity score weighting only. 2Propensity score weighting and demographics, comorbidities, and concomitant medications as regressors and stratifiers. 3Primary composite outcome includes myocardial infarction, cardiac arrest, coronary artery bypass, coronary angioplasty, heart failure, stroke, inpatient death. Bold values indicate statistically significant results.

| Reference drug | Hazard ratios for DPP-4 inhibitors use | Acute myocardial infarction | Stroke | Heart failure |
|----------------|--------------------------------------|-----------------------------|--------|-------------|
| Sulfonylureas   | HR (95% CI)\(^a\)                    | 0.77 [0.69, 0.93]           | 0.81 [0.56, 1.07] | 1.03 [0.84, 1.34] | 0.59 [0.35, 0.89] |
|                | aHR (95% CI)\(^a\)                  | 0.84 [0.74, 0.93]           | 0.95 [0.72, 1.24] | 1.08 [0.89, 1.35] | 0.71 [0.55, 0.90] |
| Metformin      | HR (95% CI)\(^a\)                    | 0.98 [0.87, 1.08]           | 1.13 [0.92, 1.33] | 1.12 [0.98, 1.34] | 0.94 [0.62, 1.41] |
|                | aHR (95% CI)\(^a\)                  | 1.07 [0.98, 1.16]           | 1.32 [0.96, 1.81] | 1.16 [0.98, 1.42] | 1.19 [0.81, 1.83] |

Table 4. Hazard ratios for the association between DPP-4 inhibitor use and primary composite outcome, stratified by cumulative exposure. Primary composite outcome includes myocardial infarction, cardiac arrest, coronary artery bypass, coronary angioplasty, heart failure, stroke, inpatient death. *Propensity score weighting and demographics, comorbidities, and concomitant medications as regressors and stratifiers. Bold values indicate statistically significant results.

| Reference drug | Hazard ratios for DPP-4 inhibitors use | Nref | NDPP-4i | Hazard ratio [95% CI] |
|----------------|--------------------------------------|------|---------|----------------------|
| Sulfonylureas   | aHR\(^a\)                            | 17,481 | 9146 | 0.84 [0.74, 0.93] |
|                | Cumulative exposure < 6 months        | 9838  | 4943   | 0.85 [0.71, 0.93] |
|                | Cumulative exposure 6–12 months       | 4326  | 2580   | 0.76 [0.70, 0.93] |
|                | Cumulative exposure 12–18 months      | 1665  | 848    | 0.80 [0.90, 1.11] |
|                | Cumulative exposure > 18 months       | 1652  | 775    | 0.98 [0.92, 1.11] |
| Metformin      | aHR\(^a\)                            | 88,596 | 9146 | 1.07 [0.98, 1.16] |
|                | Cumulative exposure < 6 months        | 47,534 | 4943 | 1.10 [0.85, 1.14] |
|                | Cumulative exposure 6–12 months       | 23,869 | 2580 | 1.05 [0.92, 1.11] |
|                | Cumulative exposure 12–18 months      | 8713  | 848    | 0.97 [0.90, 1.12] |
|                | Cumulative exposure > 18 months       | 8480  | 775    | 1.16 [0.87, 1.13] |

Results of analyses excluding individuals with acute renal failure (sulfonylureas: aHR 0.83, 95% CI 0.17–0.86; metformin: aHR 1.08, 95% CI 1.01–1.19) showed qualitatively similar results to those including these individuals (sulfonylureas: aHR 0.84, 95% CI 0.74–0.93; metformin: aHR 1.07, 95% CI 0.98–1.16). Similarly, results were not sensitive to changes in the latency period after the last dose of exposure for the primary analysis results, nor did they differ substantively when lagging the latency period by 7-days or 30-days (Additional File 1, Supplementary Appendix 4).

Discussion
In this longitudinal study of commercial claims data for individuals with diabetes, having also cardiovascular disease and renal impairment, new use of DPP-4i was associated with a lower risk for MACE compared to sulfonylureas, and a comparable risk compared to metformin. This difference was most evident in the first year of use. New use of DPP-4i was not shown to be associated with the following individual components of MACE: heart failure, stroke, or acute myocardial infarction. These results contribute to the body of evidence examining the association of DPP-4i and MACE in individuals at higher risk for cardiovascular events.

Overall, our results corroborated those of previous clinical trials and observational studies of cardiovascular safety of DPP-4i. Our results showed no increased risk for MACE with the use of DPP-4i similar to the conclusions in the three completed clinical trials comparing DPP-4i and placebo\(^{12,13,18}\). However, unlike these trials our study population consisted of a high-risk group of individuals with established cardiovascular disease and renal impairment in a real-world setting with multiple comorbidities and concomitant medications. Additionally, our study allowed us to compare DPP-4i to other available therapies, namely sulfonylureas and metformin, showing a decreased risk for MACE when compared to sulfonylureas.
Three notable observational studies provide additional context to our results. First, a 2015 administrative claims study using the Taiwan National Health Insurance Research Database of individuals with diabetes and chronic kidney disease admitted for acute myocardial infarction also showed no increased risk for MACE when comparing individuals on sitagliptin to those not on sitagliptin. Expanding on this approach, our study design included individuals with a multitude of cardiovascular conditions at baseline, making our results more generalizable to high-risk individuals with diabetes. The second study of 127,555 individuals which used the Italian Nationwide OsMed Health-DB database showed decreased risk in hospitalization for heart failure risk associated with DPP-4i compared to sulfonylurea (HR 0.78; 95% CI 0.62–0.97). This suggests that the increased risk for MACE in our comparison might be driven by differences in risk for heart failure. This is partially evident in the statistically significant unadjusted hazard ratio for heart failure as a secondary outcome (Table 3). Finally, a study of Medicare patients with diabetes compared cardiovascular risk of DPP-4i to sulfonylurea and thiazolidinediones and found no increased risk in individuals over age 65. While our study was limited to a patient population under age 65, this study shows that our results are similar to those found in a study of older patients.

A major strength of our study was the identification of a high-risk cohort through the use of a large administrative database. The results of the SAVOR TIMI-53 trial showed that renal impairment was independently associated with adverse cardiovascular outcomes, even after adjusting for cardiovascular risk factors. As such, our decision to restrict the study cohort to individuals with diabetes, comorbid with cardiovascular disease and renal impairment allowed us to hone in on a patient population identified by FDA as more appropriate for the study of this drug-event association. In a 2008 Guidance for Industry, regulators noted that such patients are often excluded from pre-approval clinical trials, and recommended studies of the association between oral antihyperglycemic agents and MACE should include high-risk patients.

Our study also had several limitations. First, we included individuals using multiple therapies of interest (e.g., both DPP-4i and sulfonylureas) at baseline. While this may have led to potential misclassification, their exclusion would have resulted in a much smaller and less generalizable sample. Second, our new-user design did not account for time-varying hazards for MACE, although the cumulative exposure stratified analysis suggested that the differences in risk of MACE between DPP-4i and sulfonylureas were most evident during the first year of exposure. Third, our analyses are subject to potential informative censoring due to a change in exposure status. However, our results were insensitive to the lagged latency period after the last day of exposure, and since most events occurred within the first year of exposure, extending the latency period past 30-days would have been unlikely to have changed our results and could have potentially led to misclassification. Finally, due to limitations in the administrative claims data, we were unable to assess mortality outside of the inpatient setting as a component of the primary composite outcome or as a competing risk. However, it is unlikely that death outside of the inpatient setting would occur more frequently in one exposure group than another in this cohort, potentially leading to biased results.

Our results provide evidence of decreased risk for MACE when comparing DPP-4i versus sulfonylurea in a commercially insured patient population with diabetes, comorbid with cardiovascular disease and renal impairment in the United States. Additionally, we found that there was no difference in risk of MACE for these individuals when comparing DPP-4i and metformin. Further studies are needed to determine differences in risk for individual components of the composite outcome, particularly heart failure. Finally, the decreased risk seen with DPP-4i use compared to sulfonylurea is more likely due to cardiovascular risk associated with sulfonylurea rather than protective effects of DPP-4i, as there was no difference in effect between DPP-4i and metformin, which carries little cardiovascular risk.

### Methods

#### Study design and data source.
We conducted a population-based, retrospective cohort study with a new-user design using data from IBM MarketScan Commercial Claims and Encounters from January 2010 through December 2015. MarketScan houses linked paid claims and encounter data from approximately 350 payers, covering more than 25 million individuals annually. The data consists of de-identified individual-level healthcare utilization data including demographic characteristics and information on inpatient and outpatient medical services and pharmacy claims issued.

#### Cohort derivation.
We identified individuals with type 2 diabetes as those with at least one prescription for an oral antihyperglycemic agent and either hemoglobin A1c greater than 6.5% twice, fasting glucose greater than 126 mg/dL twice on different days, random glucose > 200 mg/dL twice on different days, one inpatient diagnosis of diabetes (ICD-9(10) codes: 250x (E11.9), 357.2 (E11.42), 366.41 (E11.36), 362.01–362.07 (E11.3*)) or outpatient diagnosis for diabetes (ICD-9(10) codes: 250x (E11.9), 366.41 (E11.36), 362.01–362.07 (E11.3*)) twice on different days. We included individuals if they received at least one prescription for an FDA-approved DPP-4i, sulfonylurea, or metformin (Additional File 1, Supplementary Appendix 1). We assigned an index date as the date of first filled prescription for one of these products. The baseline period was defined as the six-month period preceding this. We used International Classification of Diseases, 9th and 10th Revisions (ICD-9/10) codes from inpatient and outpatient records to identify those with a history of cardiovascular disease and/or renal impairment (chronic kidney disease or acute renal failure) for inclusion into the study.

We excluded individuals based on: (1) no continuous medical or pharmacy enrollment in the six-month baseline period; (2) below the age of 35; (3) insulin use at baseline; (4) end-stage renal disease in the baseline period; (5) less than 12-weeks of exposure to index treatment; or (6) treatment with oral antihyperglycemic agents in the 6-month baseline period. We followed individuals until the first of the following occurrences: (1) 14-days after the last date of exposure; (2) switch to or addition of anti-hyperglycemic treatment that was not
Definition of exposure. We assigned individuals to an exposure group based on their first prescription of DPP-4i, sulfonylurea, or metformin. In the event of multiple drug class prescriptions on the index date, individuals were counted in all relevant exposure groups. Due to the large number of individuals on multiple antihyperglycemic agents, particularly metformin in combination with other drug classes, dropping these individuals would have resulted in a considerably smaller sample size, reducing statistical power as well as the generalizability of our results.

Definition of outcome. We defined our primary composite outcome as the first of any of the following events: myocardial infarction, cardiac arrest, coronary artery bypass graft, coronary angioplasty, heart failure, and stroke. We did not examine all-cause mortality outside of the inpatient setting as an outcome, because information on these deaths were not available for our dataset. Our secondary outcomes were acute myocardial infarction, stroke, and heart failure. We used validated algorithms to identify myocardial infarction23, cardiac arrest24, coronary artery bypass graft23, coronary angioplasty23, heart failure25, and stroke26.

Definition of covariates. We used the peer-reviewed literature19,27–29, clinical guidelines30–34 and expert opinion35 in order to identify key covariates of interest. We assessed possible confounding due to patient age and sex; cardiovascular risk factors including hypertension and hypercholesterolemia, diabetic complications, diabetic severity at baseline as measured by the Adjusted Diabetes Comorbidities Severity Index (aDCSI)36–38, and other common comorbidities such as cerebrovascular disease, ischemic heart disease, asthma, and rheumatoid arthritis. Comorbidities coded under ICD-9 were identified using the Clinical Classification Software (CCS) developed by the Agency for Healthcare Research and Quality39. Additionally, we used National Drug Codes to assess for possible confounding due to concomitant medications including statins, angiotensin converting enzymes (ACE) inhibitors, platelet aggregation inhibitor, calcium channel blockers, angiotensin II receptor blockers, beta-blockers, diuretics, nitrates, and inhaled corticosteroids.

Propensity score. We used the Toolkit for Weighting and Analysis of Nonequivalent Groups (twang) package developed by the RAND Corporation40 to compute the propensity scores and associated weights used in the analysis to balance the covariates between exposure groups. To do so, we identified all available covariates without an association to the exposure that were associated to the primary outcome in order to increase precision41 (Additional File 1, Supplementary Appendix 2). Next, we used generalized boosted regression models to optimize the selection of covariates for the propensity score calculation, allowing for propensity scores estimation in the presence of multiple exposure groups. We used the standardized mean difference (SMD) to measure balance of covariates before and after weighting. Propensity score weighting reduced the SMD from a maximum of 0.23 to less than 0.01. We used the average treatment effect on the treated (ATT) propensity score weights to estimate the treatment effect of DPP-4i.

Statistical analysis. We identified a baseline period of 6 months prior to the first filled prescription of the exposure group. Using chi-square statistics for categorical covariates and the Kruskal–Wallis test for continuous covariates, we compared differences at baseline between new users of DPP-4i, sulfonylurea, and metformin. We used exact Poisson tests to compare absolute differences in incidence rates for the primary and secondary outcomes for new users of DPP-4i compared to those of sulfonylurea and metformin. Additionally, we checked for differences in duration of exposure to treatment distributions using a Kolmogorov–Smirnov test. Using the propensity score weights described above, we calculated weighted crude and adjusted Cox proportional hazards for the association between new use of DPP-4i and the primary and secondary outcomes compared to new use of sulfonylureas and metformin. For the adjusted hazard ratios, we included indicators for age, sex, baseline comorbidities, and concomitant medications. Additionally, we checked for covariates that violated the proportional hazards assumption by plotting the scaled Schoenfeld residuals. These covariates were added to the Cox proportional hazards model as stratiﬁcation variables. We also used spline intervals every 6 months to address potential violations of the proportional hazard assumption. Finally, we included an indicator variable for individuals included in multiple exposure groups and an indicator variable for cumulative exposure, which was defined as the number of days exposed to the exposure group drug. We stratified the analysis by the cumulative exposure variable. All analyses were conducted in R, version 3.3.3.

Sensitivity analyses. To check the robustness of our results, we first assessed possible sensitivity of our results to the latency of the period after drug discontinuation. We followed individuals for 14-, 7-, and 30-days after the last day of exposure to drug. Next, we recalculated the primary analysis without individuals with acute renal failure to determine whether our results were sensitive to the inclusion of those individuals. Finally, we compared the association of DPP-4i and the primary outcome between cumulative exposure strata and sex strata to assess possible effect modification.

The study was exempted from review by a Johns Hopkins Institutional Review Board, and all research was performed in accordance with relevant guidelines/regulations. Our study utilized commercial claims data through IBM Marketscan Research Databases. As such, we did not obtain informed consent. This was in accordance with local laws and regulations, and aligned with the terms of use from IBM Marketscan Research Databases.
Ethics approval. The study was exempted from review by a Johns Hopkins Institutional Review Board.

Data availability
The data that support the availability of this study are available from IBM MarketScan Research Databases but restrictions apply to the availability of these data, which were used under license for the current study, and so are not publicly available. Data are however available from the authors upon reasonable request and with permission of IBM MarketScan Research Databases.

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Author contributions
S.B., G.A., J.S., M.M. and S.E. contributed to the development of the research question and the writing of this manuscript. S.B., J.W., H.C. and O.M. contributed to the derivation of the cohort for this study. All authors contributed to the analysis plan and interpretation of results. Analysis was conducted by S.B. and reviewed by all authors.

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Competing interests
Dr. Alexander is Chair of FDA’s Peripheral and Central Nervous System Advisory Committee, has served as a paid advisor to IQVIA, is a co-founding Principal and equity holder in Monument Analytics, a health care consultancy whose clients include the life sciences industry as well as plaintiffs in opioid litigation, and serves on OptumRx’s National P&T Committee. This arrangement has been reviewed and approved by Johns Hopkins University in accordance with its conflict of interest policies. All other authors have no conflict of interest.

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