Abstract Overweight and obesity are associated with the metabolic syndrome (MetS). We studied the association of weight change over three consecutive 5-year periods with the number of MetS components in people aged 20–59 years. 5735 participants from the Doetinchem Cohort Study were included. Weight was measured in round 1 and at each 5-year interval follow-up (round 2, 3 and 4). Weight change was defined as the absolute weight change between two consecutive measurements. The number of MetS components (assessed in round 2, 3 and 4) was based on the presence of the following components of the MetS: central obesity, raised blood pressure, reduced high density lipoprotein cholesterol and elevated glucose. Associations of weight change and the number of components of the MetS were analyzed with Generalized Estimating Equations for Poisson regression, stratified for 10-year age groups. For each age group, 1 kg weight gain was positively associated with the number of components of the MetS, independent of sex and measurement round. The association was stronger in 30–39 years (adjusted rate ratio: 1.044; 95%CI: 1.040–1.049) and smaller in older age groups. Compared to stable weight (≤2.5 kg and < 2.5 kg), weight loss (≤–2.5 kg) and weight gain (≥2.5 kg) was associated with a lower and higher rate ratio respectively, for the number of components of the MetS. Our results support the independent association of weight change with the number of MetS components with a more pronounced association in younger people.

Keywords Cardiovascular disease · Longitudinal · Metabolic syndrome · Weight change

Abbreviations
ATP III · Adult treatment panel III
BMI · Body mass index
GEE · Generalized estimating equations
HDL-cholesterol · High density lipoprotein cholesterol
MetS · Metabolic syndrome
NHG · Dutch College of General Practitioners
TNO · Netherlands Organization of Applied Scientific Research

Introduction

The prevalence of overweight (body mass index (BMI): 25–30 kg/m²) and obesity (BMI ≥ 30 kg/m²) is increasing rapidly [1]. In the Netherlands, the prevalence of overweight increased from one-third of the adult population in 1981 to about half of the adult population in 2004, and the prevalence of obesity doubled to 10% [2]. Obesity is strongly associated with cardiovascular diseases, diabetes mellitus [3], and the metabolic syndrome (MetS) [4–6].

The MetS predisposes an individual to cardiovascular morbidity and mortality [8] and diabetes mellitus [9]. Although several definitions of the MetS exist [10], the main components of the MetS are hypertension, hyperglycemia, central obesity, and dyslipidemia (raised triglycerides and/or reduced levels of high density lipoprotein cholesterol, reduced high density lipoprotein cholesterol, and raised triglycerides).
cholesterol; HDL-cholesterol) [11]. The etiology of the MetS is not completely understood, and both abdominal obesity [12, 13] and insulin resistance [7, 14] have been suggested as predominant underlying risk factors.

In the 1990s, about one million (14%) Dutch adults were reported to have the MetS [15]. In the United States, the prevalence of the MetS has increased in the nineties and is likely to rise in the future, since the MetS is related to BMI [4]. Although high weight and BMI are associated with the MetS and its components [4–6], the patterns of weight change in relation to the occurrence of the components of the MetS are less well understood. In several retrospective and prospective studies, weight gain or an increasing BMI was associated with the development of the MetS [16–20]. However, those studies were limited to young adults only [19], were performed in males only [16], or had a relatively small sample size [17]. Furthermore, some were partly based upon self-reported weight (recall) [16, 20], which tends to underestimate real weight [21].

Weight gain tends to be larger in younger individuals compared to older people [22], whereas the prevalence of the MetS increases with age [23]. Therefore, it is of importance to study the relationship between weight change and the components of the MetS in a broad age range. Hence, the aim of the present study is to describe the relationship between weight change over 5-year periods and the number of components of the MetS in a population-based longitudinal cohort study, for a priori stratified age groups. This relationship will be studied in the Doetinchem Cohort Study, a population-based sample of 12,439 men and women aged 20–59 years at baseline, who have been followed for a period of 16 years.

**Methods**

Participants and study design

The study design is previously described elsewhere [24]. In brief, the Doetinchem Cohort Study is a prospective longitudinal study performed among inhabitants aged 20–59 years of Doetinchem, a town in a rural area in the Netherlands. The aim of the Doetinchem Cohort Study is to study the impact of (changes in) lifestyle factors and biological risk factors on aspects of health. Participants were randomly selected from an age- and sex-stratified sample from the civil registries of the town of Doetinchem. At baseline (1987–1991: round 1), 20155 persons were invited to visit the municipal health centre to participate in the ‘Monitoring Project on Cardiovascular Disease Risk Factors’. The participation rate was 62% (12439 participants). In 1993–1998 (round 2), a random sample of 7769 persons was re-invited. Not all participants could be invited due to logistic reasons. In 1998–2002 (round 3) and 2003–2007 (round 4), those who were invited in round 2 were invited again, excluding those who emigrated, actively withdrew from the study or died. Interval between the first and second round was 6 years and interval between each follow-up was 5 years. However, we refer to 5 year intervals for all consecutive measurements for simplicity. The response rates for all follow-up measurements varied between 75% and 80%, resulting in 6118, 4917 and 4521 participants for round 2, 3 and 4, respectively. In the present analyses, we included the participants with data on at least one set of weight change over 5 years and the corresponding number of MetS components. For example, one set is the weight change between round 1 and 2 and the number of MetS components in round 2 (see Fig. 1). Those who reported to have (had) cancer (N = 401) were excluded from the analysis because of the potential weight loss associated with this disease. In addition, those who were pregnant during a measurement round were excluded for that specific measurement round (N = 16), resulting in 5735 participants to be included in the analysis. We certify that all applicable institutional and governmental regulations concerning the ethical use of human volunteers were followed during this research. The study was approved according to the guidelines of the Helsinki Declaration by the external Medical Ethics Committee of the Netherlands Organization of Applied Scientific Research (TNO).

**Fig. 1** Overview of analyses of 5 year weight change and number of components of the MetS. Weight change is measured between round 1 and 2 and associated with the number of components of the MetS in round 2 (white arrows and white box). Weight change between round 2 and 3 was associated with the number of components of the MetS in round 3 (gray arrows and gray box). Weight change between round 3 and 4 was associated with the number of components of the MetS in round 4 (black arrows and black box). These three sets of data were stacked and jointly analyzed with GEE.
Measurements

The study consisted of a standardized measurement of biological factors by trained personnel during a visit to the municipal health service. These measurements included anthropometric measurements, blood pressure measurements and blood sampling. Demographic characteristics, health-related information, use of medication, and lifestyle factors were examined by means of a questionnaire, which was completed at home before the physical examination.

Body weight was measured in all four rounds to the nearest 100 g on calibrated scales by trained staff, with participants wearing light indoor clothing with empty pockets and without shoes. To adjust for light indoor clothing, 1 kg was subtracted from the measured weight. Weight change was defined as the absolute change of weight in a person between two consecutive measurement rounds. Height was measured to the nearest 0.5 cm.

Central obesity was measured by means of the waist circumference, which was measured to the nearest 0.5 cm at the middle of the distance between the lowest rib and the iliac crest, with participants in standing position wearing underclothing. Systolic and diastolic blood pressure were measured twice, with the participant in sitting position. Systolic pressure was recorded at the appearance of sounds (first-phase Korotkoff) and diastolic blood pressure was recorded at the disappearance of sounds (fifth-phase Korotkoff). The mean value of two measurements was used in the analysis. Non-fasting levels of HDL-cholesterol were determined at the Lipid Reference Laboratory of the Erasmus Medical Centre in Rotterdam, using standardized enzymatic methods. Non-fasting plasma glucose level was measured with the hexokinase method [25].

Educational level was assessed as the highest level of education reached during all measurements and was classified into three categories: low (intermediate secondary education or less), medium (intermediate vocational or higher secondary education) and high (higher vocational education or university). Smoking status was categorized as being a smoker, an ex-smoker or a never-smoker. Alcohol use was based on the number of glasses of alcohol per day and was categorized in none (0 glasses), light (≤1), moderate (1–3 for men and 1–2 for women) and heavy (≥3 for men and ≥2 for women) alcohol intake. Physical activity level was assessed with an extensive questionnaire and dichotomized in those who did and did not comply with being moderately active for at least 3.5 h per week.

Number of MetS components

The cut-off values of the components of the MetS were defined according to the 2005 version of the Adult Treatment Panel (ATP) III [26], with some modifications for glucose. Central obesity was defined as a waist circumference of ≥102 cm in men and ≥88 cm in women, respectively. High blood pressure was defined as a systolic blood pressure of ≥130 mmHg, or a diastolic blood pressure of ≥85 mmHg, or use of antihypertensive medication. Reduced HDL-cholesterol was defined as <1.03 mmol/l in men and <1.3 mmol/l in women or use of cholesterol lowering medication. Random plasma glucose measurements were used, because no fasting glucose was available. According to the recommendations of the Dutch College of General Practitioners (NHG) guidelines [27], a random glucose of ≥7.8 mmol/l was considered as elevated glucose. In addition, those who used medication for diabetes were also categorized as having elevated glucose.

Blood pressure and HDL-cholesterol were determined in each round. Waist circumference and glucose were measured from the second round onwards. Glucose was not determined in the first year of round 2, therefore the number of components of the MetS could not be calculated for the first year of round 2. Triglycerides were not available. The number of components of the MetS was calculated for round 2, 3 and 4, and its maximum is four.

Statistical analyses

First, baseline characteristics of the study population were described. Second, weight and the components of the MetS were described for each round and per 10-year age group (20–29 years, 30–39 years, 40–49 years and 50–59 years), based upon baseline age.

To study the relationship between 5-year changes in weight and the number of MetS components after these 5 year periods, we included all available sets of weight change and number of MetS components (see Fig. 1). To adjust for the correlation between repeated measurements in a person, we used generalized estimating equations (GEE). GEE analyses can also deal with unequal numbers of measurements and with unequally spaced observations. Based upon the correlations of the metabolic parameters over time, we chose the exchangeable correlation structure as the working correlation structure. The number of components of the MetS was analyzed using Poisson regression, which provides a rate ratio.

The relationship between weight change and the number of components of the MetS was analyzed in two ways: (1) weight change over 5 years as a continuous determinant, and (2) weight change over 5 years as a categorical determinant; i.e. three groups: stable weight (less than 2.5 kg weight loss and less than 2.5 kg weight gain (>−2.5 kg and <2.5 kg)), weight loss (at least 2.5 kg weight loss (≤−2.5 kg)), and weight gain (at least 2.5 kg weight gain (≥2.5 kg)), with stable weight as the reference group.
2.5 kg cut-off point reflects a weight gain of approximately 0.5 kg per year. For both analyses, the first model was adjusted for sex, round, and weight in the preceding round (Model A), and the second model was further adjusted for educational level, smoking status, alcohol use, physical activity and height (Model B). Both analyses were a priori stratified for 10-year age groups, based on age in round 1.

Effect modification of sex, weight in the preceding round, and educational level was examined by testing the significance of the interaction term with weight change in the model. Effect modifiers are only mentioned when they were consistently found in at least three of the four age groups.

Because weight and waist circumference are correlated and waist circumference is one of the components of the MetS, the GEE analyses were repeated after removing central obesity from the components of the MetS in additional analyses.

Finally, the association of weight change with the continuous values of the individual metabolic components waist circumference, systolic and diastolic blood pressure, HDL-cholesterol, and glucose were studied with linear GEE analyses. Effect modification and confounding were studied similar to the other GEE analyses.

Statistical significance was set at $p < 0.05$. All analyses were carried out in SAS 9.1 for Windows (SAS Institute Inc., Cary, NC, USA) and the GEE analyses were performed with the procedure GENMOD.

Results

Table 1 presents the baseline characteristics of the study population. A total of 2735 men (47.7%) and 3000 women (52.3%) were included in the analysis. For each round and 10-year age group, weight and MetS parameters are described in Table 2. The average weight increased over time. Waist circumference and blood pressure tended to worsen over the age groups and during follow-up, whereas the change in HDL-cholesterol and glucose over the age groups and over time was less clear. The number of components of the MetS increased during follow-up. For instance, the prevalence of three or more components was 10.1% in the second round compared to 17.8% in the fourth round. The most common MetS component for all age groups and measurements combined was hypertension (55%), followed by central obesity (39%), reduced HDL-cholesterol (31%) and elevated random glucose (5%) (data not shown). The average number of MetS components (range 0–4) in round 2 was 0.75, 0.91, 1.24, and 1.61 for the age groups 20–29 years, 30–39 years, 40–49 years, and 50–59 years, respectively.

Table 1 Baseline characteristics of the study sample (Data are mean (SD) or number (%))

| Characteristics                  | Total (N = 5735) |
|----------------------------------|-----------------|
| Age (years)                      | 39.5 (10.1)     |
| Sex                              |                 |
| Men                              | 2735 (47.7%)    |
| Women                            | 3000 (52.3%)    |
| Educational level                |                 |
| Low                              | 2779 (48.5%)    |
| Moderate                         | 1687 (29.4%)    |
| High                             | 1266 (22.1%)    |
| Smoking status                   |                 |
| Current smoker                   | 1947 (34.0%)    |
| Former smoker                    | 1649 (28.8%)    |
| Never smoker                     | 2134 (37.2%)    |
| Physical activitya               |                 |
| Active                           | 2557 (45.8%)    |
| Not active                       | 2111 (45.2%)    |

* Measured at T2

In Fig. 2, the percentage individuals with weight loss ($\leq -2.5$ kg), stable weight ($>-2.5$ kg and $< 2.5$ kg) and weight gain ($\geq 2.5$ kg) over 5 years are presented per age group. The percentage of persons with weight gain became smaller with increasing age. The participants aged 20–29 years did not only have the highest percentage weight gainers (53%), but also the highest absolute increase in weight (6.8 kg ± 4.0) in the weight gain group. Furthermore, average weight change over 5 years is larger in younger age groups (20–29 years: $3.1 \pm 5.4$ kg) than in older age groups (50–59 years: $1.0 \pm 4.2$ kg).

Table 3 (Model A) presents the regression coefficients (adjusted for sex, measurement round and weight in the preceding round) for the association of 1 kg weight change over 5 years with the number of components of the MetS. The a priori stratification for age groups was appropriate because significant interactions with age were found. For all age groups, 1 kg weight change was positively associated with the number of components, with rate ratios ranging from 1.025 (95% CI: 1.021–1.029) in the 50–59 years old to 1.042 (95% CI: 1.038–1.046) in the 30–39 years old. The latter rate ratio can be interpreted as follows: an increase in weight of 1 kg in 5 years was associated with 1.042 (95% CI: 1.038–1.046) times more components of the MetS. Adjustment for educational level, smoking status, alcohol use, physical activity and height (Model B) did not materially change the relationship between weight change and the number of components of the MetS. Significant effect modification was present for weight change and weight in the preceding round: the association of weight change with the...
The number of components of the MetS was slightly less strong in persons with a higher weight in the preceding round.

Table 4 presents the association of weight loss and weight gain over a period of 5 years with the number of components of the MetS. For all age groups, weight loss...
was associated with a decreased number of components and weight gain was associated with an increased number of components, compared to the stable weight group. Not surprisingly, these associations were the strongest in those aged 20–29 years because their absolute weight loss and weight gain was the largest. In persons aged 20–29 years, weight gain was associated with 1.305 (95% CI: 1.212–1.405) times the number of components of those who remained relatively stable in weight. In persons aged 50–59 years, ≥2.5 kg weight gain was related to only 1.141 (95% CI: 1.101–1.182) times the number of components of those in the stable weight group. Similar differences between the age groups were found for the association of weight loss and stable weight. In the fully adjusted model (Model B), the effect of weight loss and gain remained present and became slightly more pronounced. There was significant effect modification between weight change category and weight in the preceding round for the weight gain group compared to the weight stable group, but not for the weight loss group. The strength of the association of weight gain with the number of MetS components was less in people with higher baseline weight.

Additional analysis

Weight and waist circumference are correlated and waist circumference is one of the MetS components, thus part of the dependent variable. Therefore, we repeated the analyses by excluding central obesity from the MetS components, resulting in a maximum of three MetS components. The association of 1 kg weight change in a 5-year period with the number of MetS components was weaker in the analysis without central obesity than in the analysis including central obesity as part of the MetS components, but still statistically significant for all age groups (rate ratios ranging from 1.015 to 1.033), indicating that the association of weight change with the number of MetS components does not depend solely on its association with waist circumference. A similar pattern of weaker but significant associations was observed for the weight gain/weight loss analyses.

Finally, we studied the associations of 1 kg weight change over 5 years with the individual metabolic components with longitudinal linear regression analysis. One

### Table 3

| Age       | Model A Rate ratio | 95% CI      | Model B Rate ratio | 95% CI      |
|-----------|--------------------|-------------|--------------------|-------------|
| 20–29 years | 1.038              | 1.031–1.045 | 1.038              | 1.030–1.046 |
| 30–39 years | 1.042              | 1.038–1.046 | 1.044              | 1.040–1.049 |
| 40–49 years | 1.031              | 1.027–1.034 | 1.033              | 1.029–1.037 |
| 50–59 years | 1.025              | 1.021–1.029 | 1.025              | 1.021–1.030 |

Table presents rate ratios from longitudinal Poisson regression. Model A is adjusted for: sex, measurement round and weight in the preceding round. Model B is additionally adjusted for: educational level, smoking status, alcohol use, physical activity and height

### Table 4

| Weight loss (≤−2.5 kg)                                                                 |
|-----------------------------------------|-----------------|-----------------|
| Model A Rate ratio | 95% CI      | Model B Rate ratio | 95% CI      |
| 20–29 years              | 0.773          | 0.685–0.872     | 0.768          | 0.680–0.867     |
| 30–39 years              | 0.790          | 0.738–0.845     | 0.783          | 0.729–0.840     |
| 40–49 years              | 0.809          | 0.768–0.853     | 0.800          | 0.758–0.844     |
| 50–59 years              | 0.888          | 0.843–0.934     | 0.873          | 0.828–0.920     |

| Weight gain (≥2.5 kg)                                                                 |
|-----------------------------------------|-----------------|-----------------|
| Model A Rate ratio | 95% CI      | Model B Rate ratio | 95% CI      |
| 20–29 years              | 1.305          | 1.212–1.405     | 1.318          | 1.224–1.418     |
| 30–39 years              | 1.282          | 1.234–1.333     | 1.310          | 1.258–1.363     |
| 40–49 years              | 1.208          | 1.169–1.249     | 1.224          | 1.183–1.267     |
| 50–59 years              | 1.141          | 1.101–1.182     | 1.132          | 1.091–1.175     |

Table presents rate ratios from Poisson regression. Reference group is the stable weight group (≥−2.5 kg and <2.5 kg). Model A is adjusted for: sex, measurement round and weight in the preceding round. Model B is additionally adjusted for: educational level, smoking status, alcohol use, physical activity and height
kg weight change was related to an increase in waist circumference ($\beta$: 0.838 cm; 95% CI: 0.821–0.856), systolic blood pressure ($\beta$: 0.432 mmHg; 95% CI: 0.379–0.486), diastolic blood pressure ($\beta$: 0.355 mmHg; 95% CI: 0.318–0.392) and glucose ($\beta$: 0.014 mmol; 95% CI: 0.009–0.020) and a decrease in HDL-cholesterol ($\beta$: −0.011 mmol; 95% CI: −0.012 to −0.011) in the fully adjusted model (model B). No consistent effect modification was observed for the individual metabolic components.

**Discussion**

In this prospective cohort study, weight change over a 5-year period was associated with increased number of components of the MetS, independent from sex, measurement round and other cardiovascular risk factors. Compared to those who remained relatively stable in weight, weight gain (≥2.5 kg) was related to an increase in the number of components of the MetS, whereas weight loss (≤−2.5 kg) was related to a decrease in the number of components. The association of weight change and the number of MetS components was stronger in younger than in older people. Furthermore, there was negative effect modification with weight in the preceding round, indicating that the association of weight change and the number of MetS components in those already high in weight was slightly weaker.

As weight gain is an important public health concern, more insight into the effect of weight change on metabolic components is needed. Because weight change and the development of MetS components varies with age [22, 23], we studied the effect of weight change on the number of MetS components for different age groups. In the present study, in all adult age groups, weight gain predisposed an individual to increased number of MetS components, compared to those who maintained a stable weight. Similar associations were observed by Hillier et al. [18], who showed in French individuals aged 30–64 years that larger weight gain was associated with higher rates of both developing and maintaining the MetS than smaller weight gain. In addition, in the CARDIA study, a prospective study in persons aged 18–30 years at baseline, a stable BMI over time was associated with a smaller incidence of the MetS compared to a rise in BMI over time, regardless of the baseline BMI [19]. Furthermore, in the longitudinal study of Ferreira et al. [17], it was found that persons with the MetS at age 36 had a more marked increase of BMI from the age of 21 onwards.

In our study, weight gain was positively related with the individual MetS components waist circumference, systolic and diastolic blood pressure, and glucose level and negatively with HDL-cholesterol. This is in accordance with the findings of several other longitudinal studies in adults with different duration of follow-up (range: 2–11 years) [18, 28–31]. Weight gain was related to unfavorable changes in waist circumference [18], blood pressure, hypertension [18, 28–31], and lipid profile [18, 30, 31]. For glucose level, Norman et al. [31] found a positive association between weight change and fasting glucose level. Czernichow et al. [29] did not observe this association, possibly due to the short follow-up time of 2 years in this study. Similar to our results, Hillier et al. [18] found that weight change was most strongly related to waist circumference out of all individual metabolic components.

It is well known that components of the MetS tend to worsen with age [19] and that the prevalence of the MetS increases with age [32, 33]. Our study is in line with these findings, because the number of components of the MetS increased with age. In our analyses with weight change as a continuous variable, those aged 30–39 years showed the largest rate ratio for the association of weight change and number of MetS components. Since absolute weight gain was larger in younger adults, it might be important to prevent weight gain in younger adults. In the weight loss/weight gain analyses, the higher rate ratio observed in young adults is probably largely the result of the higher absolute weight change in young adults.

In our study we found that those who lost weight had a reduction in the number of components compared to the participants who had a stable weight. The favorable effect of weight loss on the metabolic risk profile has also been shown in weight loss intervention studies in obese adults [34]. However, weight loss and weight loss maintenance are hard to achieve [35]. Recent studies demonstrated an association of repeated weight loss [36] and weight fluctuation [20] with the MetS. Therefore, weight stabilization may be a more important public health goal [19]. Weight gain prevention in adults can be accomplished by preventing a positive energy balance by stimulating physical activity and reducing energy intake, and by focusing on the environment to make the healthy choice the easiest choice [37].

Because no standard definition of weight loss, stable weight and weight gain exists [38], we set the cut-off value of weight change on 2.5 kg. Sensitivity analyses with cut-off values of 1 and 4 kg confirmed our findings that the association of weight change with the number of components of the MetS was stronger in younger age groups (data not shown).

What distinguishes the present study from earlier ones is that the relationship between weight change and the MetS is quantified in a broad age group (20–59 years) with a follow-up of 16 years. Another strength is that weight change was not self-reported, but measured by trained personnel. Further strengths are its large sample size, the
prospective design, and the availability of up to four measurements per individual. The response rate of the follow-up measurements was also relatively high ($\geq 75\%$ from the second round onwards). Finally, by using GEE analysis all data are effectively used.

One of the limitations is selective loss to follow-up. Compared to the participants included in this study, those who were lost to follow-up were more often lower educated, current smokers and of younger age. In addition, baseline BMI and baseline blood pressure were higher in those who were lost to follow-up, whereas baseline HDL-cholesterol was lower. No difference was found for sex, weight and alcohol consumption at baseline (data not shown). Our study population is therefore not representative for the Dutch population, and the results might not be generalizable to young, low educated persons with a relative adverse cardiovascular risk profile. Another important limitation is the lack of information about triglycerides. Therefore, it was not possible to determine the relationship between weight change and the MetS. Instead, our study included four metabolic components, derived from the 2005 version of the ATP III definition of the MetS [26] and looked at the number of MetS components. In addition, glucose level was measured non-fasting. This might be less accurate since food intake affects the level of blood glucose [39]. We compensated for this by using a higher cut-off value reflecting fasting glucose levels, as reported in Rutten et al. [27].

While it seems that certain components of the MetS are prone to cluster, the exact components and cut-off values of the MetS differ over the existing definitions [12]. Furthermore, the number of MetS components in round 2 is somewhat larger in older age groups than in younger people. The different rate ratio’s observed for the age groups might also reflect a ceiling effect. However, as the number of metabolic syndrome components at baseline in the oldest age group is not extremely high, we feel that the ceiling effect does not completely account for the observed difference in rate ratio’s between age groups.

With the dichotomization of the components information is lost. Analyses with the continuous values of the individual components showed associations in the expected direction for each separate component and strengthened our findings.

In sum, weight gain is positively associated with the number of components of the MetS, whereas weight loss is negatively associated with this number. Our study shows that the effect of weight change on the number of MetS components is different for various age groups. Prevention of weight gain might be relevant in younger adults in particular, for whom the association of weight change with the number of MetS components is the strongest and the weight gain is the largest.

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