Folate Intake and the Risk of Breast Cancer: A Dose-Response Meta-Analysis of Prospective Studies

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Abstract

Background: Previous observational studies regarding the existence of an association between folate intake and the risk of breast cancer have been inconsistent. This study aimed to summarize the evidence regarding this relationship using a dose-response meta-analytic approach.

Methodology and Principal Findings: We performed a systematic search of the PubMed, EmBase, and Cochrane Library databases to identify studies published through June 2013. Only prospective observational studies that reported breast cancer effect estimates with 95% confidence intervals (CIs) for more than 2 folate intake categories were included. We excluded traditional case-control studies because of possible bias from various confounding factors. Overall, we included 14 prospective studies that reported data on 677,858 individuals. Folate intake had little effect on the breast cancer risk (relative risk (RR) for highest versus lowest category = 0.97; 95% CI, 0.90–1.05; P = 0.451). Dose-response meta-analysis also suggested that a 100 μg/day increase in folate intake had no significant effect on the risk of breast cancer (RR = 0.99; 95% CI, 0.98–1.01; P = 0.361). Furthermore, we used restricted cubic splines to evaluate the nonlinear relationship between folate intake and the risk of breast cancer, and discovered a potential J-shaped correlation between folate intake and breast cancer risk (P = 0.007) and revealed that a daily folate intake of 200–320 μg was associated with a lower breast cancer risk; however, the breast cancer risk increased significantly with a daily folate intake >400 μg.

Conclusion/Significance: Our study revealed that folate intake had little or no effect on the risk of breast cancer; moreover, a dose-response meta-analysis suggested a J-shaped association between folate intake and breast cancer.

Introduction

Given its role as a modulator of DNA synthesis, repair, and methylation, folate was hypothesized to reduce the risk of breast cancer [1]. The importance of these processes in cell growth and development led to investigations of the consequences of either low or high folate intake on cancer development [2]. The results of several studies suggested that increased folate intake was associated with an increased prostate cancer risk [3] but significantly reduced risks of esophageal, stomach, pancreatic [4], and colorectal cancers [5]. However, data regarding the subsequent effects of folate intake on breast cancer are limited and inconclusive.

A meta-analysis [6] based on 9 prospective studies and 14 case-control studies suggested no clear association between the folate intake or blood folate levels and the risk of breast cancer; however, folate appeared to significantly counteract the increased risk of breast cancer associated with moderate or high levels of alcohol consumption. It is particularly important to clarify the optimal daily folate intake level with respect to the general population, as this level has yet not been definitively determined. Furthermore, additional unanswered questions remain, including whether associations differ according to follow-up duration, and alcohol intake.

Folate intake has been studied in numerous prospective studies of primary breast cancer prevention and, in the present study, we attempted a large-scale examination of these available prospective studies to update the results and determine the association between folate intake and the risk of breast cancer. Furthermore, we performed a dose-response meta-analysis to quantitatively determine the optimal folate intake level with respect to the general population.

Methods

Data Sources, Search Strategy, and Selection Criteria

This review was conducted and reported according to the Preferred Reporting Items for Systematic Reviews and Meta-Analysis Statement [7] issued in 2009 (Checklist S1).
Any prospective study that examined the relationship between folate intake and breast cancer was eligible for inclusion in our study, and no restrictions were placed on the publication language or status (published, in press, or in progress). We searched the PubMed, EmBase, and Cochrane Library electronic databases for articles published through June 2013 using the following search terms ("folate" OR "folic acid") AND ("cancer" OR "neoplasm" OR "carcinoma") AND ("cohort" OR "cohort studies" OR "nest case-control studies"). We also conducted manual searches of the reference lists from all relevant original and review articles to identify additional eligible studies. The medical subject headings, methods, patient populations, designs, exposures, and outcome variables of these articles were used to identify the relevant studies. Two of the authors (HFG and YFZ) conducted this literature search independently, according to a standardized approach. Any inconsistencies between the 2 authors were settled by the primary author (YHZ) until a consensus was reached. Studies were eligible for inclusion if the following criteria were met: (1) the study had a prospective design (prospective cohort or nested prospective case-control study); (2) the study investigated the association between folate intake and the risk of breast cancer; and (3) the authors reported the effect estimates (risk ratio [RR], hazard ratio [HR], or odds ratio [OR]) and 95% confidence intervals (CIs) for comparisons of highest and lowest category folate intake and included ≥2 folate intake categories. We excluded all case-control studies because various confounding factors could have biased the results.

Data Collection and Quality Assessment

The collected data included the first author’s or study group’s name, publication year, country, study design, folate exposure assessment, sample size, age at baseline, follow-up duration, effect estimate and 95% CI, comparison categories, and covariates in the fully adjusted model. We also extracted the numbers of cases/persons or person-years, the effects of the different exposure categories, and the 95% CIs. For studies that reported several multivariable adjusted-effect estimates, we selected the effect estimate that had been maximally adjusted with respect to potential confounders.

The highly comprehensive Newcastle-Ottawa Scale (NOS) [8], which has been partially validated for quality evaluations of observational studies in meta-analyses, was used to evaluate the methodological quality [9]. The NOS is based on the following 3 subscales: selection (4 items), comparability (1 item), and outcome (3 items). A “star system” (range, 0–9) was developed for assessment purposes (Table S1). The data extraction and quality assessments were conducted independently by 2 authors (HFG and LZ). The information was examined and adjudicated independently by an additional author (YHZ) with reference to the original studies.

Statistical Analysis

We examined the relationship between folate intake and the risk of breast cancer on the basis of the effect estimates (RR or HR) and 95% CI published in each study. We first used the random-effects model [10,11] to calculate the summary RRs and 95% CIs for highest versus lowest category folate intake levels. Second, we transformed category-specific risk estimates into RR estimates associated with an increase in folate intake of 100 μg/day by using the method of generalized least squares for trend estimation [12]. These estimates were calculated by assuming a linear relationship between the natural logarithm of RR and increasing folate intake. The value assigned to each folate category was the mid-point for closed categories and the median for open categories (assuming a normal distribution for folate intake). We combined the RRs for each 100 μg/day increase in folate intake by using the results of a random-effect meta-analysis [10]. Third, We conducted a dose-response random-effects meta-analysis of the correlated natural logs of the RRs or HRs across all folate intake categories [12,13]. To derive the dose-response curve, we modeled folate using...
Table 1. Baseline characteristic of studies included in the systematic review and meta-analysis.

| Study                  | Country          | Study design | Assessment of exposure | Sample size | Age at baseline | Effect estimate | Comparison categories | Follow-up (year) | Covariates in fully adjusted model                                                                                                                                                                                                                                                                                                                                 |
|------------------------|------------------|--------------|------------------------|-------------|-----------------|-----------------|----------------------|-----------------|---------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| Zhang S 1999 [18–19]   | US               | Cohort       | FFQ                    | 88818       | 30–55           | RR              | >600 vs 150–299 ug/d| 16.0            | Age, length of follow-up, total energy intake, parity, age at first birth, history of BC in mother or a sister, history of benign BC, alcohol intake, BMI at age 18 years, weight change from age 18 years, height, age at menopause, HRT use, and beta-carotene                                                                                      |
| Rohan TE 2000 [20]     | Canada           | Cohort       | FFQ                    | 56837       | NG              | RR              | >354.28 vs < 224.78 ug/d | 8.0            | Energy intake, age, age at menarche, number of live births, menopausal status, family history of BC in a first-degree relative, practice of breast self-examination, alcohol consumption, randomization group, and study center                                                                                                           |
| Sellers TA 2001 [21–22]| US               | Cohort       | FFQ                    | 34387       | 62.0            | RR              | >294 vs < 172 ug/d   | 12.0           | Age, education, family history of BC, age at menarche, age at menopause, oral contraceptive use, HRT, parity, age at first birth, BMI, waist-to-hip ratio, height, BMI at age 18, alcohol, smoking, PA, and other B vitamins                                                                                                                                  |
| Feigelson HS 2003 [23–24]| US              | Cohort       | FFQ                    | 66561       | 40–87           | RR              | >918.9 vs <192.2 ug/d| 14.0           | Age, includes ethanol, dietary folate, methionine, vitamin B vitamin, race, education, first-degree family history of BC, history of breast lump, mammographic history, HRT use, parity and age at first birth, age at menopause, age at menarche, PA, BMI, adult weight gain, and energy                                                                 |
| Stolzenberg-Solomon RZ 2006 [27]| US           | Cohort       | FFQ                    | 25400       | 55–74           | HR              | 400 vs 200 ug/d    | 13.8            | Total energy and folate intake and fitted as linear variable in Cox’s proportional hazard model with age as time metric                                                                                                                                                                                                 |

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| Study               | Country     | Study design | Assessment of exposure | Sample size | Age at baseline | Effect estimate | Comparison categories | Follow-up (year) | Covariates in fully adjusted model                                                                                                                                 |
|--------------------|-------------|--------------|------------------------|-------------|----------------|-----------------|---------------------|------------------|--------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| Lajous M 2006      | French      | Cohort       | FFQ                    | 62739       | 40–65          | RR              | 522 vs 296 ug/d    | 9.0              | Age, two-year follow-up period, region of residence, years of education, family BC, history of benign breast disease, age at menarche, parity, breastfeeding, years since last use of oral contraceptives, age at menopause, years of HRT use, regular mammographic evaluation, height in cm, BMI, vitamin supplement use, alcohol intake and PA. |
| Tjønneland A 2006  | Denmark     | Nest case control | FFQ                  | 24697       | 50–64          | RR              | >400 vs <300 ug/d | 4.7              | Vitamin C, total energy, school education, BMI, parous/nulliparous and number of births, age at birth of first child, history of benign breast tumour surgery.                                                      |
| Ericson U 2007     | Sweden      | Cohort       | FFQ                    | 11699       | >50            | HR              | 456 vs 160 ug/d    | 9.5              | Weight, height, leisure-time PA, household work, smoking, socioeconomic status, age at menopause, HRT, and alcohol intake.                                                                                                                     |
| Larsson SC 2008    | Sweden      | Cohort       | FFQ                    | 61433       | 40–70          | RR              | >277 vs <200 ug/d | 17.4             | Age, education, body mass index, height, parity, age at first birth, age at menarche, age at menopause, use HRT, family history of breast cancer, history of benign breast disease, and intakes of alcohol and total energy. |
| Lin J 2008         | US          | Nest case control | FFQ                  | 28345       | >45            | RR              | 582 vs <263.9 ug/d | 11.0             | Matching variables, age, randomized treatment assignment, BMI, family history of BC in a first-degree relative, history of benign breast disease, smoking, PA, alcohol consumption, age at menarche, age at menopause, parity, and age at first birth. |
| Maruti SS 2009     | US          | Cohort       | FFQ                    | 35023       | 50–76          | RR              | 541 vs 210 ug/d    | 5.0              | Age, race, family history of BC, mammography within 2 y preceding baseline, history of breast biopsy, age at menarche, age at first birth, age at menopause, years of HRT use, BMI, height, total PA, and alcohol intake in the past year, total energy intake. |
restricted cubic splines with 3 knots at fixed distribution percentiles of 10%, 50%, and 90% [12]. This method required knowledge of the distributions of cases and persons or person-years and effect estimates (RRs or HRs) along with the variance estimates for at least 3 quantitative exposure categories.

Inter-study heterogeneity was investigated with the Q statistic, and we considered P-values < 0.10 as indicative of significant heterogeneity [14,15]. Breast cancer subgroup analyses were conducted on the basis of the country, study design, sample size, effect estimate (HR or RR), follow-up duration, adjusted alcohol intake, and alcohol intake. We also performed a sensitivity analysis by removing each individual study from the meta-analysis.

Several methods were used to evaluate the potential publication bias. Funnel plots for the risk of breast cancer were visually inspected. The Egger [16] and Begg [17] tests were also used to statistically assess the publication bias with respect to the breast cancer incidence. All reported P values were 2-sided, and P values < 0.05 were considered statistically significant for all included studies. STATA software (version 12.0; Stata Corporation, College Station, TX, USA) was used to perform the statistical analyses.

Results

The results of the study-selection process are shown in Figure 1. We identified 640 articles during our initial electronic search, of which 492 were excluded as duplicates or irrelevant studies. A total of 158 potentially eligible studies were thus selected. After a detailed evaluation, 14 cohorts from 19 studies [18–36] were selected for the final meta-analysis. A manual search of the reference lists of these studies did not yield any new eligible studies. The general characteristics of the included studies are presented in Table 1.

Of the 14 included cohorts (for a total of 677,858 individuals), 12 cohorts were from 16 prospective cohort studies [18–28,31,32,34,36] and 2 cohorts were from 3 nested case-control studies [29,30,33]. The participant follow-up period ranged 4.7–17.4 years and the number of individuals per study ranged 11,699–88,818. Seven cohorts [18,19,21–24,27,33–35] were based in the US, 5 [25,26,28–32] in Europe, 1 [20] in Canada, and 1 [36] in China. The study quality was assessed according to the NOS [8] (Table S1). Herein, we considered a study with a score ≥ 7 to be high quality. Overall, 3 cohorts [20,23,24,31] had scores of 9, 9 cohorts [18,19,21,22,25,30,32,34,36] had scores of 8, 1 cohort [33] had a score of 7, and 1 cohort [35] had a score of 6.

After pooling the included studies, the summary RR revealed that folate intake was not associated with breast cancer (RR for highest versus lowest category = 0.97; 95% CI, 0.90–1.05; P = 0.451; Figure 2); however, evidence of potentially significant heterogeneity was noted (I² = 57.5%; P = 0.004). Consequently, a sensitivity analysis was conducted and after sequentially excluding each study from the pooled analysis, the conclusion was not found to be affected by the exclusion of any specific study. The dose-response meta-analysis findings did not suggest any association between the risk of breast cancer and a 100 mg/day increase in folate intake (RR, 0.99; 95% CI, 0.98–1.01; P = 0.361; Figure 3), with substantial heterogeneity across studies (I² = 66.2%; P < 0.001).

All studies were included in the dose-response meta-analysis of the relationship between folate intake and the risk of breast cancer. As shown in Figure 4 and by the P-value for nonlinearity (P = 0.007), we discovered evidence of a nonlinear relationship between folate intake and the risk of breast cancer. A daily folate intake of 200–320 mg was associated with reduced the risk of
A heterogeneity assessment of this analysis yielded a P-value < 0.10. Accordingly, we conducted subgroup analyses to minimize heterogeneity among the included studies. Overall, we noted that folate intake level was associated with a reduced breast cancer risk if the patients had a daily alcohol intake < 10 g (RR for highest versus lowest category = 0.64; 95% CI: 0.43–0.97). Conversely, folate intake was associated with an increased risk of breast cancer when nested case-control studies were included (RR for highest versus lowest category = 1.23; 95% CI: 1.01–1.50; Figure 5). Furthermore, Subgroup analysis revealed that a 100 µg/day increment in folate intake was also associated with increased breast cancer risk (RR = 1.04; 95% CI: 1.01–1.07; Figure 5) when nested case-control studies were included. No other significant differences were identified with respect to the effects of increased folate intake in association with additional factors.

Figure 2. Relative risk estimates of breast cancer for highest versus lowest folate intake category.

doi:10.1371/journal.pone.0100044.g002

Figure 3. Dose-response meta-analysis for per 100 µg/day increment in folate intake for breast cancer.

doi:10.1371/journal.pone.0100044.g003
A review of the funnel plots did not rule out the possibility of publication bias with respect to breast cancer (Figure 6). However, the Egger [16] and Begg [17] test results did not show any evidence of publication bias, P = 0.936 and P = 0.913 for highest versus lowest category folate intake, respectively, and P = 0.576 and P = 0.274 for a 100 \( \text{mg} \)/day increment in folate intake, respectively.

**Discussion**

The current meta-analysis evaluated prospective studies to explore all possible correlations between folate intake and the risk of breast cancer. This large quantitative study included 677,858 individuals from 12 prospective cohort studies and 2 nested case-control studies with a broad population range. The findings from our current meta-analysis suggest that there are no effects of increased folate intake on the incidence of breast cancer.

A previous meta-analysis [6] of observational studies (cohort studies, nested case-control studies, and case-control studies) suggested that folate intake or blood folate levels had no effect on the risk of breast cancer and that adequate folate intake significantly reduced the risk of breast cancer in individuals with moderate or high levels of alcohol consumption. However, the various confounding factors present in case-control studies could lend bias to the results. Furthermore, the cutoff points for the folate intake categories differed among the studies. Finally, we also reported the association between folate intake and breast cancer in some specific subsets. Another meta-analysis of randomized controlled trials that was published by the B-Vitamin Treatment Trialists’ Collaboration [37] suggested that folic acid supplementation did not affect the risk of breast cancer. An inherent limitation of the earlier meta-analysis was that the included trials had been designed to evaluate the effects of folic acid supplementation on cardiovascular or other outcomes rather than cancer-related outcomes; additionally, the results had been derived from very few cases and should thus have been considered preliminary results. Furthermore, the follow-up duration was insufficiently long to demonstrate a clinical benefit and thus always yielded broad confidence intervals (i.e., no statistically significant difference). Thus far, no study has confirmed the association between folate intake and the risk of breast cancer. Therefore, we conducted a dose-response meta-analysis of the existing prospective studies in order to identify the optimal folate intake dose.

Most of our findings agreed with those of a recently published large cohort study conducted in the US [18]; that prospective study included 88,818 individuals and found that participants who consumed >600 \( \text{mg} \)/day of folate per day exhibited a 7% decrease in the risk of breast cancer when compared with individuals who consumed 150–299 \( \text{mg} \)/day, although this decrease was not statistically significant. Furthermore, adequate folate intake might reduce the excess risk of breast cancer associated with alcohol consumption. Rohan et al. [20] also suggested that increased folate intake did not affect the overall risk of breast cancer but significantly reduced the risk of breast cancer in participants with a daily alcohol consumption level >14 g. Our current study also indicated that increased folate intake level had no significant effect on the overall risk of breast cancer but that folate might play an important role with respect to the risk of breast cancer in participants with high alcohol consumption levels. Possible reasons for these findings might include the different folate intake category cutoff points among the studies as well as data collection methods that only compared the breast cancer risks at highest and lowest category folate intake levels. Furthermore, because folate could potentially promote tumor cell growth, high blood folate levels might be associated with an increased risk of breast cancer [38,39].

No significant difference was observed between increased folate intake and the risk of breast cancer. However, several studies that were included in our meta-analysis reported inconsistent results. Lajous et al. [29] indicated that folate intake level was associated with a reduced risk of breast cancer (highest versus lowest

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**Figure 4. Dose-response relations between folate intake and relative risks of breast cancer.**

doi:10.1371/journal.pone.0100044.g004
A. Subgroup analysis for highest versus lowest folate intake category for breast cancer

| Country       | RR (95% CI)       |
|---------------|-------------------|
| US            | 1.01 (0.91, 1.11) |
| Other         | 0.92 (0.81, 1.06) |
| Study design  |                  |
| Cohort study  | 0.95 (0.87, 1.02) |
| Nest case–contrib | 1.23 (1.01, 1.50) |
| Sample Size   |                  |
| >500000       | 0.94 (0.87, 1.03) |
| <500000       | 1.01 (0.86, 1.18) |
| Effect estimate|                |
| RR            | 0.97 (0.89, 1.07) |
| HR            | 0.94 (0.79, 1.13) |
| Duration of follow-up |       |
| more than 10 years | 0.98 (0.91, 1.06) |
| less than 10 years | 0.96 (0.82, 1.12) |
| Adjusted alcohol intake |         |
| yes           | 0.94 (0.85, 1.03) |
| no            | 1.04 (0.90, 1.19) |

B. Subgroup analysis for per 100 ug/day increment in folate for breast cancer.

| Country       | RR (95% CI)       |
|---------------|-------------------|
| US            | 1.00 (0.99, 1.01) |
| Other         | 0.98 (0.95, 1.01) |
| Study design  |                  |
| Cohort study  | 0.99 (0.98, 1.00) |
| Nest case–control study | 1.04 (1.01, 1.07) |
| Sample Size   |                  |
| >500000       | 0.99 (0.97, 1.00) |
| <500000       | 1.00 (0.98, 1.02) |
| Effect estimate|                |
| RR            | 0.99 (0.98, 1.01) |
| HR            | 1.00 (0.97, 1.02) |
| Duration of follow-up |       |
| more than 10 year | 0.99 (0.98, 1.01) |
| less than 10 years | 1.00 (0.97, 1.02) |
| Adjusted alcohol intake |         |
| yes           | 0.99 (0.97, 1.00) |
| no            | 1.01 (0.99, 1.03) |

Figure 5. Subgroup analysis for breast cancer.
doi:10.1371/journal.pone.0100044.g005
Furthermore, Ericson et al. [31] also suggested that folate intake was associated with reduced the risk of breast cancer by 44% (highest versus lowest category). However, those 2 cohorts had low power with which to detect the interaction between folate and alcohol intake and, moreover, most of the women in those cohorts consumed alcohol whereas few used folate supplements. These 2 factors could possibly explain why those studies found that folate exhibited an overall protective effect against breast cancer.

In restricted cubic splines analysis, we identified a potential nonlinear relationship between folate intake and breast cancer suggestive of an association between a daily folate intake of 200–320 μg and a lower breast cancer risk; however, a daily folate intake >400 μg appeared to significantly increase the risk of breast cancer. A possible explanation for this finding is that folate normally acts as a modulator of DNA synthesis, repair, and methylation [1] but at a daily folate intake >400 μg levels, folate might affect endothelial function and support cell growth [40].

A subgroup analysis suggested that increased folate intake was associated with reduced the risk of breast cancer in patients with a daily alcohol consumption level >10 g. Several studies [18–20] included in our analysis reported results consistent with this finding. Furthermore, folate intake was associated with increased

A Funnel plot for high versus low folate intake for breast cancer

![Funnel plot with pseudo 95% confidence limits](image1)

Egger test, P=0.936  Begg’s test, P=0.913

B. Funnel plot for per 100 μg/day increment in folate for breast cancer.

![Funnel plot with pseudo 95% confidence limits](image2)

Egger test, P=0.576  Begg’s test, P=0.274

Figure 6. Funnel plot for breast cancer.
doi:10.1371/journal.pone.0100044.g006
the risk of breast cancer when we pooled the nested case-control studies. However, this conclusion might have been unreliable because smaller cohorts were included in that subset. Therefore, we simply reported the relative result and thus provided a synthetic and comprehensive review.

Two strengths of our study should be highlighted. First, only prospective studies were included; this restriction should have eliminated selection and recall biases, which might have been of concern with regard to the retrospective case-control studies. Second, the dose-response analysis included a broad folate intake range, thus allowing an accurate assessment of the dose-response relationship between folate intake and the risk of breast cancer.

The limitations of our study were as follows: (1) publication bias is an inevitable problem in meta-analyses of published studies; (2) data regarding breast cancer in premenopausal or postmenopausal women were unavailable; and (3) the analysis used pooled data (individual data were not available), which restricted us from performing a more detailed relevant analysis and obtaining more comprehensive results.

The results of this study suggest that increased folate intake has no significant effect on the risk of breast cancer. According to our dose-response meta-analysis, a daily folate intake of 200–320 µg appeared to associate with a lower risk of breast cancer; in contrast, increased breast cancer risk was associated with a daily folate intake >400 µg/d. Future studies should focus on specific populations in order to analyze primary breast cancer prevention.

Supporting Information

Table S1 Quality scores of prospective cohort studies using Newcastle-Ottawa Scale.

Table S2 Daily intakes of folate and risk of breast cancer data in included studies.

Checklist S1 PRISMA Checklist.

Author Contributions

Conceived and designed the experiments: Y-HZ. Performed the experiments: Y-HZ H-FG W-WS A-JH Y-FZ. Analyzed the data: Y-HZ LZ. Contributed reagents/materials/analysis tools: Y-HZ. Wrote the paper: Y-HZ.

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