Mobility and social deprivation on primary care utilisation among paediatric patients with asthma

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ABSTRACT

Objective Asthma care is negatively impacted by neighbourhood social and environmental factors, and moving is associated with undesirable asthma outcomes. However, little is known about how movement into and living in areas of high deprivation relate to primary care use. We examined associations between neighbourhood characteristics, mobility and primary care utilisation of children with asthma to explore the relevance of these social factors in a primary care setting.

Design In this cohort study, we conducted negative binomial regression to examine the rates of primary care visits and annual influenza vaccination and logistic regression to study receipt of pneumococcal vaccination. All models were adjusted for patient-level covariates.

Setting We used data from community health centres in 15 OCHIN states.

Participants The sample included 23 773 children with asthma aged 3–17 across neighbourhoods with different levels of social deprivation from 2012 to 2017. We conducted negative binomial regression to examine the rates of primary care visits and annual influenza vaccination and logistic regression to study receipt of pneumococcal vaccination. All models were adjusted for patient-level covariates.

Results Clinic visit rates were higher among children living in or moving to areas with higher deprivation than those living in areas with low deprivation (rate ratio (RR) 1.09, 95% CI 1.02 to 1.17; RR 1.05, 95% CI 1.00 to 1.11). Children moving across neighbourhoods with similarly high levels of deprivation had increased RRs of influenza vaccination (RR 1.13, 95% CI 1.03 to 1.23) than those who moved but stayed in neighbourhoods of low deprivation.

Conclusions Movement into and living within areas of high deprivation is associated with more primary care use, and presumably greater opportunity to reduce undesirable asthma outcomes. These results highlight the need to attend to patient movement in primary care visits, and increase neighbourhood-targeted population management to improve equity and care for children with asthma.

INTRODUCTION

There is consistent and compelling evidence that neighbourhood characteristics affect paediatric asthma outcomes. Asthma is influenced by factors related to the physical environment (air quality, built environment), social position in the neighbourhood (race/ethnicity, immigration status) and other social environmental factors, such as those relating to neighbourhood deprivation, wealth and inequality.1–4 Research has shown that living in or moving into an impoverished neighbourhood is associated with higher risk of asthma when compared with never having lived in a poor neighbourhood.5

More frequent changes in neighbourhood may also influence asthma outcomes or asthma care utilisation. A Welsh team, using data from more than 200 000 children, found that moving within the first year of life was associated with increased risk of future emergency hospitalisation for asthma.6 Other studies from Europe and New Zealand have also found residential mobility to be associated with negative health outcomes7,8 possibly because moves can be an indication of financial insecurity or other underlying family instability.
According to the National Asthma Education and Prevention Program’s Expert Panel, one of the main goals in asthma care is asthma control, which includes reducing impairment (symptoms) and reducing risk (e.g., preventing exacerbations and reducing emergency care). Primary care clinics are essential in this endeavour, because they are where many asthma management medications are prescribed and asthma action plans are developed. Primary care clinics are also increasingly screening for a wide range of socioeconomic adversity indicators, yet screening is labour intensive and therefore often not feasible in busy clinical environments. An alternative could be to use administrative data on address changes and zip-code linked neighbourhood-level deprivation data to identify clinical populations at risk for negative outcomes. If this risk stratification is meaningful, it could help prioritise care for populations with greater needs without necessitating large-scale patient-facing screening. Address data are readily accessible by many clinics, and could be used to therefore better assess risk. However, these associations between neighbourhood-level deprivation, moving neighbourhoods and primary care utilisation in children with asthma has not been studied at a national level in the USA, though some work has studied neighbourhood-level deprivation on a smaller scale.

Community health centres (CHCs) are clinics in the USA that provide primary care to anyone regardless of insurance status or the patient’s ability to pay for care in a defined community. They serve disproportionately high numbers of low-income patients, including residents of public housing and those experiencing homelessness. These features make CHCs particularly appropriate settings in which to examine how residential mobility across neighbourhoods with varying levels of social deprivation affects routine primary care utilisation (regular visits and immunisations) in children with asthma. We hypothesised that children with asthma who receive care in CHCs would have higher care needs and therefore higher primary care utilisation if they (1) moved from areas with little deprivation to areas with more deprivation and (2) had higher number of residential moves within areas of high deprivation.

METHODS

Data and inclusion criteria

Our analysis uses OCHIN (formerly Oregon Community Health Information Network until other states joined) data on children with asthma from 15 states, aged 3–17 who had ≥1 ambulatory visit in study clinics between 1 January 2012 and 31 December 2017 and had documented geocoded addresses. Guidelines encourage regular clinical contact in children with asthma, and it is especially important to vaccinate children with asthma for other respiratory illnesses. Study participant flow based on inclusion criteria is reported in online supplemental figure 1.

Outcomes

We chose the outcomes of office visits and immunisations because they apply to all children with asthma regardless of asthma severity or current symptoms. Specifically, the outcome variables were (1) rate of primary care visits (the National Asthma Education and Prevention Program Expert Panel recommends that children with asthma be seen twice/year at minimum); (2) annual rate of influenza immunisation, as children who have asthma are also at an increased risk of complications from influenza, which may require hospitalisation, and (3) a binary indicator of whether the child had received ≥1 pneumococcal vaccination, as the pneumococcal immunisation is recommended for children with asthma, who may be more at risk for invasive pneumococcal disease.

Asthma diagnosis was determined by having ≥1 International Classification of Disease (ICD)-9 codes of 493* or ICD-10 codes of J45* documented at an visit during the study period. We identified influenza and pneumococcal immunisations using the Centers for Disease Control and Prevention (CDC)’s National Center of Immunization and Respiratory Diseases vaccine administered codes, which were reviewed and cleaned by analysts, and reviewed again for final categorisation by a practicing clinician.

Independent variables

The main independent variable was a categorical indicator of residential mobility during the study period with five groups: (1) all moves within high deprivation neighbourhoods (‘always high’); (2) all moves within low deprivation neighbourhoods (‘always low’); (3) a move from a highly deprived neighbourhood to one with low deprivation (‘high to low’); (4) a move from a neighbourhood with low deprivation to one with high deprivation (‘low to high’); and (5) multiple moves (>2 moves) across neighbourhoods with at least one differing deprivation level (‘varied’). Children who did not move at all either fell into the always high deprivation group or always low deprivation group. Addresses were obtained from the electronic health record (EHR); these are collected as a matter of routine care. The current address recorded in the EHR is a reliable measure of exposure to neighbourhood level deprivation. To measure social deprivation in an area, we used the Social Deprivation Index (SDI) of the census tract in which the patient resided at the time. The SDI is a validated composite measure of neighbourhood-level social determinants including income, employment, education, housing and transportation, and can be used to represent the overall socioeconomic conditions in a geographical area. The index was created by the Robert Graham Center for the year 2015, using a subset of 2011–2015 American Community Survey 5-year estimates. The SDI score (a unitless range of 0–100) was split at the median (based on this sample, median=83) to create the ‘low deprivation’ and ‘high deprivation’ categories. Low deprivation included...
values between 0 and 83, high deprivation included values between 83 and 100. The models also include the number of moves experienced by each child (categorised as 0, 1–2, and ≥3 moves). 26

Patient-level characteristics added as covariates included: age at first visit (categorised in years as 3–5, 6–11, 12–17), sex (female or male), categorical indicator of rate of visits per year in the models analysing vaccination (<1, 1–2, 3–4, ≥5), primary insurance type during the study period recorded at each visit (never insured, some private insurance, some public insurance, some public and private insurance), income as percent of US federal poverty level (FPL), and race/ethnicity (Latinx, non-Hispanic white, non-Hispanic black). While we generally use ‘Latino/a/x’ because it can be preferred in populations similar to our study population, the actual ethnicity information collected by clinics is ‘Hispanic’ and ‘non-Hispanic white’. We also included body mass index (BMI) over time, as asthma severity and control may be influenced by obesity. 27 28 BMI was calculated from EHR data from clinic visits and categorised as never overweight/obese, overweight/obese at some clinic visits, or overweight/obese at every clinic visit during the study period. Overweight was defined as having a BMI ≥85th percentile to <95th percentile for one’s respective age and sex, while obesity was defined as a BMI ≥95th percentile, consistent with CDC standards. 29 BMI and BMI percentile values between 0 and 83, high deprivation included values between 83 and 100. The models also include the number of moves experienced by each child (categorised as 0, 1–2, and ≥3 moves). 26

Results

The sample was composed of 23,773 children with asthma who received primary care in CHCs. Table 1 shows patient characteristics. The sample included Latinx (46.5%), non-Hispanic white (27.7%), and non-Hispanic black (25.8%) children. Slightly more than half the population was male (55.5%) and many patients first visit in the study period was at age 3–5 years (38.7%) or age 6–11 years (39.2%). Most patients had some public insurance (84.5%) and 1–2 ambulatory visits per year (45.4%). In this sample, 33.8% of patients were overweight/obese at every visit and 36.9% were never overweight/obese. The majority of children (62.1%) lived in a household reporting income <138% FPL at every ambulatory visit. The annual rate of influenza vaccination was 0.72 vaccines per year, and 25.1% had received a pneumococcal vaccination within the study period. The majority of children had 0 residential moves within the study period (65.6%), while 31.3% had 1–2 residential moves within the study period. The largest proportion of patients were in the always high (50.2%) or always low (36.2%) deprivation mobility groups. 6.2% of children were in the high-to-low deprivation group, and 4.7% were in the low-to-high group (table 1).

Children living in the always high deprivation group and in the low-to-high deprivation group had higher adjusted RRs of clinical visits than those in the always low group (RR 1.09, 95% CI 1.02 to 1.17 and RR 1.05, 95% CI 1.00 to 1.11, respectively) (table 2). The annual influenza immunisation rate was also higher in the always high deprivation group (RR 1.13, 95% CI 1.03 to 1.23) compared with the always low group (table 2). Similar ORs of pneumococcal vaccination were seen between all groups (table 3). Full model results are reported in online supplemental tables 1 and 2.

Sensitivity analysis: stratified by number of residential moves

Children with no recorded residential moves who lived in the always high SDI areas had increased RRs of visits compared with children with no recorded moves living in the always low areas (table 4). All other SDI mobility groups had similar RRs of visits, regardless of whether they had 1–2 residential moves or ≥3 residential moves within the study period.

Influenza vaccine rates were higher in children with no moves in the always high group compared with the always low group. Influenza vaccine rates were also higher in the always high group compared with the always low group in those with 1–2 moves, and in the always high and varied groups compared with the reference group for children with ≥3 moves (table 4). Odds of pneumococcal vaccination were increased among those with ≥3 moves in the high-to-low, low-to-high, and varied groups compared with those in the always low reference category (table 5).
We acknowledge the large confidence intervals due to small sample size in table 5, and interpret these results with caution.

**DISCUSSION**

This study used EHR data from a multistate network of clinics linked with administrative and location data to examine the effects of social deprivation and residential mobility on three indicators of routine primary care utilisation in a CHC cohort of children with asthma, preschool age and older. In an era of increasing attention to and collection of place-based social determinants of health data, it is crucial to know how neighbourhood features and residential mobility are associated with healthcare utilisation.

| Table 1 | Patient characteristics overall and in children with asthma (N=23 773) |
|---------|---------------------------------------------------------------------|
| **Race/ethnicity** | |
| Non-Hispanic white | 6581 (27.7) |
| Latinx | 11 067 (46.5) |
| Non-Hispanic black | 6125 (25.8) |
| **Age at first visit** | |
| 3–5 years | 9201 (38.7) |
| 6–11 years | 9319 (39.2) |
| 12–17 years | 5253 (22.1) |
| **Sex** | |
| Female | 10 578 (44.5) |
| Male | 13 195 (55.5) |
| **Insurance type** | |
| Never insured | 429 (1.8) |
| Some public insurance | 20 887 (84.5) |
| Some private insurance | 1829 (7.7) |
| Some public and private | 1428 (6.0) |
| **Average visits per year** | |
| Less than 1 | 2626 (11.1) |
| 1–2 | 10 799 (45.4) |
| 3–4 | 5627 (23.7) |
| 5 or more | 4721 (19.9) |
| **BMI category** | |
| Never overweight/obese | 8772 (36.9) |
| Sometimes overweight/obese | 6958 (29.3) |
| Always overweight/obese | 8043 (33.8) |
| **Per cent of US federal poverty level** | |
| Always ≥138 | 1498 (6.3) |
| Above and below 138 | 2628 (11.1) |
| Always <138 | 14 764 (62.1) |
| Never documented | 4883 (20.5) |
| **Influenza immunisation** | |
| Annual rate (95% CI) | 0.72 (0.71 to 0.73) |
| **Pneumococcal vaccine (ever)** | |
| No | 15 716 (74.9) |
| Yes | 5268 (25.1) |

| Table 2 | Rate ratios of visits and influenza vaccinations |
|---------|------------------------------------------------|
| **Visits (N=23 773)** | **Influenza vaccination (N=20 984)** |
| **RR (95% CI)** | **RR (95% CI)** |
| **Social deprivation mobility group** | |
| Always low | Ref | Ref |
| Always high | 1.09 (1.02 to 1.17) | 1.13 (1.03 to 1.23) |
| High to low | 0.90 (0.84 to 0.97) | 0.89 (0.83 to 0.97) |
| **No of residential moves** | |
| 0 | Ref | Ref |
| 1–2 | 0.91 (0.87 to 0.95) | 0.93 (0.89 to 0.98) |
| 3 or more | 0.85 (0.78 to 0.93) | 0.80 (0.73 to 0.89) |

| Table 3 | ORs of pneumococcal vaccination ever (within study period) (N=20 984) |
|---------|---------------------------------------------------------------------|
| **Social deprivation mobility group** | |
| Always low | Ref |
| Always high | 1.03 (0.93 to 1.15) |
| High to low | 1.03 (0.88 to 1.21) |
| Low to high | 1.15 (0.96 to 1.37) |
| Varied | 1.25 (0.98 to 1.60) |
| **No of residential moves** | |
| 0 | Ref |
| 1–2 | 1.27 (1.15 to 1.41) |
| 3 or more | 0.92 (0.70 to 1.21) |

Models adjusted for race/ethnicity, age at first visit, sex, insurance type, BMI, income. Vaccination model also adjusted for visits per year. Bold values indicate statistical significance. BMI, body mass index.
We found three significant associations worthy of further discussion. First, living or moving within/to a high social deprivation area was associated with increased utilisation of primary care and rates of influenza vaccination. This overarching pattern was supported by significant associations in the whole study population (living within or moving to a high deprivation area), as well as positive findings in the stratified analysis (higher likelihood of visits in the children with no moves in a high deprivation neighbourhood and a higher influenza vaccine rate across all move strata in high deprivation neighbourhoods.) That children with asthma in deprived neighbourhoods appear to use more primary care has several important implications. It may imply worse control of this chronic disease (though we did not limit these visits to asthma visits), or that they have increased healthcare needs in general. But while neighbourhood deprivation may signal providers and clinics that children may use more care (and require more proactive disease or health management), it also implies that CHCs may be an important point of contact for these children. Rather than less interaction with these centres, children from deprived neighbourhoods have more interaction with primary care than their less deprived peers, and CHCs could be viewed as valuable platforms for social and health interventions in these communities.

Second, moving was associated with less visits and influenza than those children who did not move. Residential mobility, in a low-income population like that which is served by CHCs, has been shown to cause

### Table 4 Rate ratios (RRs) of visits and influenza vaccinations, stratified by number of residential moves

| Mobility group         | Visits                     | Influenza vaccinations |
|------------------------|----------------------------|------------------------|
|                        | 0 moves (n=15597)          | 1–2 moves (n=7436)     | ≥3 Moves (n=740)     |
|                        | RR (95% CI)                | RR (95% CI)            | RR (95% CI)          |
| Always low             | Ref                        | Ref                    | Ref                  |
| Always high            | 1.10 (1.03 to 1.18)        | 1.04 (0.95 to 1.14)    | 1.12 (0.91 to 1.39)  |
| High to low            | N/A                        | 0.97 (0.91 to 1.04)    | 1.03 (0.83 to 1.27)  |
| Low to high            | N/A                        | 1.02 (0.95 to 1.09)    | 1.15 (0.92 to 1.44)  |
| Varied                 | N/A                        | 0.96 (0.88 to 1.06)    | 1.10 (0.94 to 1.30)  |

| Mobility group         | 0 moves (n=13544)          | 1–2 moves (n=6747)     | ≥3 moves (n=693)     |
|------------------------|----------------------------|------------------------|----------------------|
|                        | RR (95% CI)                | RR (95% CI)            | RR (95% CI)          |
| Always low             | Ref                        | Ref                    | Ref                  |
| Always high            | 1.12 (1.02 to 1.22)        | 1.15 (1.03 to 1.29)    | 1.28 (1.02 to 1.60)  |
| High to low            | N/A                        | 1.05 (0.96 to 1.14)    | 1.06 (0.81 to 1.38)  |
| Low to high            | N/A                        | 1.06 (0.97 to 1.15)    | 1.14 (0.84 to 1.54)  |
| Varied                 | N/A                        | 1.05 (0.91 to 1.20)    | 1.23 (1.01 to 1.49)  |

Models adjusted for race/ethnicity, age at first visit, sex, insurance type, BMI, income. Vaccination model also adjusted for visits per year. Bold values indicate statistical significance.

### Table 5 ORs of pneumococcal vaccination ever (within the study period) stratified by number of residential moves

| Mobility group         | 0 moves (n=13544)          | 1–2 moves* (n=6747)     | ≥3 moves* (n=693)     |
|------------------------|----------------------------|-------------------------|-----------------------|
|                        | OR (95% CI)                | OR (95% CI)             | OR (95% CI)           |
| Always low             | Ref                        | Ref                     | Ref                   |
| Always high            | 0.95 (0.83 to 1.10)        | 0.89 (0.72 to 1.11)     | 2.71 (0.91 to 8.11)   |
| High to low            | N/A                        | 0.90 (0.75 to 1.09)     | 5.84 (1.76 to 19.43)  |
| Low to high            | N/A                        | 1.01 (0.80 to 1.28)     | 6.01 (1.82 to 19.89)  |
| Varied                 | N/A                        | 1.12 (0.80 to 1.58)     | 3.24 (1.14 to 9.19)   |

Model adjusted for race/ethnicity, age at first visit, sex, insurance type, BMI, income, and visits per year. Bold values indicate statistical significance.

*Due to small sample size the insurance covariate was collapsed to ‘never insured/public insurance’ and ‘all else’.

BMI, body mass index; N/A, not available.
disruptions in health and healthcare.8–10 In this analysis, we also find evidence that such mobility is also associated with reduced routine care utilisation in children with a chronic disease. From these findings we are unable to determine if these children received sufficient or quality care, but the fact that they use less than their counterparts raises concern that this is possible, and clinical organisations and future research could consider further evaluating risk screening strategies that employ number of address changes from administrative data in order to stratify clinical risk.

Thirdly, more moves were associated with higher rates of receiving pneumococcal vaccination, as evidenced by an increased frequency of this immunisation in the whole sample with 1–2 moves, and in the stratified sample with 3+ moves across multiple neighbourhood types. While pneumococcal vaccination in our age strata of children with asthma may be routine, it may be more common in children who did not complete the entire pneumococcal series earlier than age three, or if they have had particularly severe symptoms requiring oral steroids.31 32 While we could not determine these specific factors from our data, the elevated odds of this immunisation in children with more moves does suggest disruptions to the completion of routine preventive measures. This lends further support to the idea that number of moves may be considered for future study as a risk stratification method.

This study had several limitations. All children received care at CHCs, therefore our findings may not be generalisable outside of this setting. While our data did include many clinical factors, we did not have any data on patient beliefs and perceptions regarding vaccination, important factors to consider when examining vaccination rates. It is also possible that children could have moved out of the CHC network, thus our rates may be underestimated. However, research in CHCs suggests that patients often do not leave their networks.3 3 Our data also did not have information on the reason for the move.

CONCLUSIONS

This analysis of the neighbourhood deprivation and residential mobility of children with asthma seeking care at CHCs demonstrated increased utilisation of primary care visits and influenza vaccinations in children in or who moved to high social deprivation neighbourhoods, and reduced utilisation of these services in children who changed addresses. Researchers and healthcare organisations should consider the opportunity for health and social interventions and care that this utilisation may provide. They should also consider further study of the incorporation of assessing residential mobility into clinical risk prediction in this cohort of children.

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Data availability statement Data may be obtained from a third party and are not publicly available. Data are a limited (partially deidentified) dataset from electronic health records. Data are available to investigators for a fee for those who are willing and able to partner with OCHIN on their analysis.

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REFERENCES

1 Barnett E, Casper M. A definition of “social environment”. Am J Public Health 2001;91:465a.
2 Chiterewe T, Shim JK, Barker JC, et al. How neighborhoods influence health: lessons to be learned from the application of political ecology. Health Place 2017;45:117–23.
3 Chen E, Hayen R, Le V, et al. Neighborhood social conditions, family relationships, and childhood asthma. Pediatrics 2019;144:e20183300.
4 Brewer M, Kimbro RT, Denney JT, et al. Does neighborhood social and environmental context impact racial/ethnic disparities in childhood asthma? Health Place 2017;44:86–93.
5 Beck AF, Huang B, Simmons JM, et al. Role of financial and social hardships in asthma racial disparities. Pediatrics 2014;133:431–9.
6 Martin G, Gavine A, Incley J, et al. Conceptualizing, measuring and evaluating constructs of the adolescent neighbourhood social environment: a systematic review. SSM Popul Health 2017;3:335–51.
7 Cantu P, Kim Y, Sheehan C, et al. Downward neighborhood poverty mobility during childhood is associated with child asthma: evidence from the geographic research on wellbeing (grow) survey. J Urban Health 2019;96:558–69.
8 Hutchings HA, Evans A, Barnes P, et al. Residential moving and preventable hospitalizations. Pediatrics 2016;138:e20152836.
9 Tennessen M, Telle K, Syse A. Childhood residential mobility and long-term outcomes. Acta Sociol 2016;59:113–29.

10 Darlington-Pollock F, Norman P, Lee AC, et al. To move or not to move? exploring the relationship between residential mobility, risk of cardiovascular disease and ethnicity in New Zealand. Soc Sci Med 2016;165:128–40.

11 National Heart, Lung, and Blood Institute, Expert Panel Report 3: Guidelines for the Diagnosis and Management of Asthma (EPR-3). Asthma care quick reference diagnosing and managing asthma. U.S. Department of Health and Human Services, editor. 2012.

12 Herrera C-N, Brochier A, Pellicer M, et al. Implementing social determinants of health screening at community health centers: clinician and staff perspectives. J Prim Care Community Health 2019;10:215013771988726.

13 Gottlieb L, Hessler D, Long D, et al. A randomized trial on screening for social determinants of health: the iScreen study. Pediatrics 2014;134:e1611–8.

14 Byhoff E, Garg A, Pellicer M, et al. Provider and staff feedback on screening for social and behavioral determinants of health for pediatric patients. J Am Board Fam Med 2019;32:297–306.

15 Nkoy FL, Stone BL, Knighton AJ, et al. Neighborhood deprivation and childhood asthma outcomes, accounting for insurance coverage. Hosp Pediatr 2018. doi:10.1542/hpeds.2017-0032. [Epub ahead of print: 09 01 2018].

16 Molina AL, Molina Y, Walley SC, et al. Residential instability, neighborhood deprivation, and pediatric asthma outcomes. Pediatr Pulmonol 2020;55:1340–8.

17 National Association of Community Health Centers. Community health center 2020 Chartbook.

18 Groshkopf LA, Aiyakan E, Broder KR, et al. Prevention and Control of Seasonal Influenza with Vaccines: Recommendations of the Advisory Committee on Immunization Practices - United States, 2019-20 Influenza Season. MMWR Recomm Rep 2019;68:1–21.

19 Vasilieilou E, Sheikh A, Butler C, et al. Effectiveness of influenza vaccines in asthma: a systematic review and meta-analysis. Clin Infect Dis 2017;65:1388–95.

20 Simon AE, Ahrens KA, Akinbami LJ. Influenza vaccination among US children with asthma, 2005-2013. Acad Pediatr 2016;16:68–74.

21 Policy Statement—Immunization for Streptococcus pneumoniae Infections in High-Risk Children. Pediatrics. 2014;134(6):1230–1233. Pediatrics 2015;135:945.

22 Centers for Disease Control and Prevention. IIS: HL7 standard code SetMapping CVX to vaccine groups, 2019. Available: https://www2a.cdc.gov/vaccines/iis/iisstandards/vaccines.asp?rpt=vg

23 Knighton AJ. Is a patient’s current address of record a reasonable measure of neighborhood deprivation exposure? A case for the use of point in time measures of residence in clinical care. Health Equity 2018;2:62–9.

24 Butler DC, Pettersson S, Phillips RL, et al. Measures of social deprivation that predict health care access and need within a rational area of primary care service delivery. Health Serv Res 2013;48:539–59.

25 Robert Graham Center. Social deprivation index (SDI), 2018. Available: https://www.graham-center.org/rcg/maps-data-tools/sdi/social-deprivation-index.html

26 Katz A, Chateau D, Enns JE, et al. Association of the social determinants of health with quality of primary care. Ann Fam Med 2018;16:217–24.

27 Forno E, Celedón JC. The effect of obesity, weight gain, and weight loss on asthma inception and control. Curr Opin Allergy Clin Immunol 2017;17:123–30.

28 Vijayakanthi N, Greally JM, Rastogi D. Pediatric obesity-related asthma: the role of metabolic dysregulation. Pediatrics 2016;137. doi:10.1542/peds.2015-0812. [Epub ahead of print: 08 04 2016].

29 Centers for Disease Control and Prevention. Defining childhood obesity, 2018. Available: https://www.cdc.gov/obesity/childhood/defining.html

30 Vogel M. childhoods: data and methods around reference values in pediatrics. R package version 0.6.7, 2017. Available: https://CRAN.R-project.org/package=childhoods

31 Eisenlohr CP, Chartrand EM, Barzaga MR, et al. Impact of pneumococcal vaccine response on asthma exacerbation frequency in young children. Immun Inflamm Dis 2020;8:493–6.

32 Pelton SI, Weycker D, Farkouh RA, et al. Risk of pneumococcal disease in children with chronic medical conditions in the era of pneumococcal conjugate vaccine. Clin Infect Dis 2014;59:615–23.

33 Huguet N, Kaufmann J, O’Malley J, et al. Using electronic health records in longitudinal studies: estimating patient attrition. Med Care 2020;58(Suppl 6 Suppl 1):S46–52.
Appendix Figure 1. Selection of Analytic Sample

Patients seen in ADVANCE clinics, age ≤18 in 2005-2017
N=1,131,205

Non-Hispanic Black, Latino, Non-Hispanic White, English and Spanish speakers
N=953,722

Patients aged 3-17, with ≥2 ambulatory encounters in a primary care or public health study clinic from 2012-2017
N=217,548

Patients with address data
N=155,947

Patients with asthma
N=23,773
| Social Deprivation Mobility Group | Encounters (N=23,773) | Influenza Vaccination (N=20,984) |
|----------------------------------|-----------------------|----------------------------------|
|                                  | RR (95% CI)           | RR (95% CI)                      |
| Always low                       | Ref                   | Ref                              |
| Always high                      | 1.09 (1.02-1.17)      | 1.13 (1.03-1.23)                 |
| High-to-low                      | 1.00 (0.94-1.06)      | 1.03 (0.95-1.12)                 |
| Low-to-high                      | 1.05 (1.00-1.11)      | 1.04 (0.96-1.13)                 |
| Varied                           | 1.02 (0.95-1.11)      | 1.08 (0.96-1.21)                 |
| Number of Residential Moves      |                       |                                  |
| 0                                | Ref                   | Ref                              |
| 1-2                              | 0.91 (0.87-0.95)      | 0.93 (0.89-0.98)                 |
| 3 or more                        | 0.85 (0.78-0.93)      | 0.80 (0.73-0.89)                 |
| Race/Ethnicity                   |                       |                                  |
| Non-Hispanic White               | Ref                   | Ref                              |
| Latino                           | 1.09 (1.01-1.17)      | 1.38 (1.23-1.54)                 |
| Non-Hispanic Black               | 0.85 (0.78-0.93)      | 1.27 (1.06-1.51)                 |
| Age at first encounter           |                       |                                  |
| 3-5 years                        | Ref                   | Ref                              |
| 6-11 years                       | 0.90 (0.87-0.93)      | 0.72 (0.67-0.78)                 |
| 12-17 years                      | 1.14 (1.04-1.25)      | 0.80 (0.73-0.89)                 |
| Sex                              |                       |                                  |
| Female                           | Ref                   | Ref                              |
| Male                             | 0.89 (0.87-0.92)      | 1.03 (0.99-1.06)                 |
| Insurance Type                   |                       |                                  |
| Some private insurance           | Ref                   | Ref                              |
| Never insured                    | 0.95 (0.78-1.16)      | 0.98 (0.77-1.24)                 |
| Some public and private          | 1.18 (1.10-1.27)      | 0.71 (0.61-0.82)                 |
| Some public insurance            | 1.10 (1.03-1.17)      | 0.90 (0.78-1.04)                 |
| BMI Category                     |                       |                                  |
| Never overweight/obese           | Ref                   | Ref                              |
| Sometimes overweight/obese       | 1.17 (1.14-1.21)      | 0.93 (0.88-0.98)                 |
| Always overweight/obese          | 1.09 (1.06-1.13)      | 0.97 (0.92-1.02)                 |
| Percent of US Federal Poverty Level |                   |                                  |
| Always ≥138                      | Ref                   | Ref                              |
| Above & below 138                | 1.09 (1.02-1.16)      | 0.74 (0.65-0.83)                 |
| Always <138                      | 1.06 (0.99-1.12)      | 0.86 (0.77-0.95)                 |
| Not documented                   | 0.87 (0.79-0.96)      | 1.21 (1.01-1.45)                 |
| Visits per Year                  |                       |                                  |
| Less than 1                      | N/A                   | Ref                              |
| 1 to 2                           | N/A                   | 2.14 (1.93-2.39)                 |
| 3 to 4                           | N/A                   | 3.40 (3.05-3.78)                 |
| 5 or more                        | N/A                   | 5.26 (4.61-6.00)                 |
### Appendix Table 2. Odds Ratios of Pneumococcal Vaccination Ever (within study period) (N=20,984)

| Social Deprivation Mobility Group          | OR (95%CI)          |
|-------------------------------------------|---------------------|
| Always low                                | Ref                 |
| Always high                               | 1.03 (0.93-1.15)    |
| High-to-low                               | 1.03 (0.88-1.21)    |
| Low-to-high                               | 1.15 (0.96-1.37)    |
| Varied                                    | 1.25 (0.98-1.60)    |
| Number of Residential Moves               |                     |
| 0                                         | Ref                 |
| 1-2                                       | 1.27 (1.15-1.41)    |
| 3 or more                                 | 0.92 (0.70-1.21)    |
| Race/Ethnicity                            |                     |
| Non-Hispanic White                        | Ref                 |
| Latino                                    | 1.23 (1.10-1.37)    |
| Non-Hispanic Black                        | 1.27 (1.10-1.47)    |
| Sex                                       |                     |
| Female                                    | Ref                 |
| Male                                      | 1.11 (1.03-1.19)    |
| Insurance Type                            |                     |
| Some private insurance                    | Ref                 |
| Never insured                             | 1.07 (0.75-1.52)    |
| Some public and private                   | 0.66 (0.48-0.89)    |
| Some public insurance                     | 1.41 (1.19-1.69)    |
| Visits per Year                           |                     |
| Less than 1                               | Ref                 |
| 1 to 2                                    | 1.73 (1.44-2.07)    |
| 3 to 4                                    | 2.59 (2.13-3.15)    |
| 5 or more                                 | 3.84 (3.15-4.69)    |
| BMI Category                              |                     |
| Never overweight/obese                    | Ref                 |
| Sometimes overweight/obese                | 0.80 (0.72-0.89)    |
| Always overweight/obese                   | 0.45 (0.44-0.53)    |
| Percent of US Federal Poverty Level       |                     |
| Always ≥138                               | Ref                 |
| Above & below 138                         | 0.60 (0.49-0.74)    |
| Always <138                               | 0.87 (0.72-1.04)    |
| Not documented                            | 0.92 (0.75-1.14)    |