Age, gender, functional KSS, reason for revision and type of bone defect predict functional outcome 5 years after revision total knee arthroplasty: a multivariable prediction model

Jan F. M. Verbeek1 · Gerjon Hannink2 · Koen C. Defoort1 · Ate B. Wymenga1 · Petra J. C. Heesterbeek3

Received: 17 August 2018 / Accepted: 18 January 2019 / Published online: 28 January 2019
© The Author(s) 2019

Abstract
Purpose The number of revision total knee arthroplasties (rTKA) is increasing. Unfortunately, not all patients benefit from revision surgery. The aim of this study was to develop a clinical prediction model that can be used to predict the functional outcome 5 years after rTKA.

Methods Data of patients receiving rTKA at Sint Maartenskliniek, Nijmegen, The Netherlands, from 2004 onwards were prospectively collected. Demographic and clinical variables and patient-reported outcome scores were collected and considered as potential predictors. Beneficial outcome was defined as an increase of ≥ 20 points on the functional knee society scores (fKSS) or an absolute fKSS ≥ 80 points 5 years after surgery. The prediction model was developed using backward logistic regression. Regression coefficients were converted into an easy to use prediction rule.

Results Overall, 295 rTKA patients were included of whom 157 (53%) had beneficial fKSS 5 years later. Age, gender, femoral bone defects, preoperative fKSS, and stiffness as reason for revision were included in the model. Men had a higher chance of beneficial fKSS than women (OR 1.59, 95% CI 0.91–2.78). Patients with major bone defects (OR 0.44, 95% CI 0.22–0.85), higher age (IQR OR 0.39, 95% CI 0.26–0.58), higher preoperative fKSS (IQR OR 0.42, 95% CI 0.30–0.59), and severe stiffness (OR 0.48, 95% CI 0.20–1.15) had a lower chance of successful outcome. The model’s AUC was 0.76, 95% CI 0.70–0.81.

Conclusion Easily determinable characteristics of patients who need rTKA can be used to predict future functional outcome. Young men with low preoperative fKSS without severe stiffness are more likely to achieve a beneficial outcome.

Level of evidence IV.

Keywords Total knee arthroplasty · Revision · Knee Society Score · Functional outcome · Clinical prediction model

Introduction

Worldwide, the number of total knee arthroplasties (TKA) is increasing. In the United States, currently more than 700,000 TKA are carried out each year [15]. Likewise, the annual number of TKA revisions has increased from 38,000 in 2005 to 55,000 in 2010, and is expected to grow even further sixfold in the year 2030 [15–18].

The aim of a revision TKA (rTKA) is to provide stable fixation, reduce pain, and improve functional outcome.
rTKA is technically a more difficult operation than primary TKA, and is associated with higher risks of complications (28–49%) [2, 30]. Some patients, however, do not benefit from rTKA or even experience a worse outcome, with major negative impact on their quality of life [28, 35]. To overcome these problems, patients should be clearly informed about what they could expect from this operation. Tools have been developed that predict failure of primary TKA [26], but no tool is available to manage patient’s expectations regarding rTKA.

A number of patient-related factors have been identified that play an important role in predicting rTKA outcome, such as age, race, body mass index, and diabetic status [9, 10, 22, 23, 26]. In addition, severe stiffness (formerly referred to as arthrofibrosis) as main reason for revision resulted in a worse outcome compared to other reasons for revision like instability, infection and malposition [33]. Prediction of functional outcome after rTKA might provide insights enabling better management of patient’s expectations. Based on the literature, we hypothesized that predictors for the functional outcome 5 years after rTKA would include age, gender, preoperative patient-reported outcome scores, reason for revision, and type of bone defect. The purpose of this study was to develop a risk prediction model to best predicting functional outcome 5 years after rTKA.

Materials and methods

Characteristics, surgical details, and outcomes of all rTKA patients treated at the Sint Maartenskliniek, Nijmegen, the Netherlands, were collected in a prospective database from 2004 onwards. We selected patients with a total system revision treated with a condylar stemmed implant (patella could be resurfaced or not, based on surgeon’s decision), and who had at least 5 years of follow-up after rTKA. Excluded were patients with a hinged arthroplasty, revision of unicondylar arthroplasty, patients with relatively small intervention revisions, such as isolated polyethylene exchange and patellar resurfacing, and patients with shorter follow-up than 5 years.

Surgical technique

All rTKA operations were performed by five orthopaedic surgeons with at least 10 years of experience in the field of revision knee arthroplasty (approximately 25 revisions per year). The surgical procedures used the standard medial parapatellar approach [1]. The cementing or hybrid fixation technique was used for the revision system, and all revision prostheses in this study were posterior stabilized or of the constrained type, including the Genesis II© or Legion© revision system (Smith & Nephew, Inc, Memphis, TN, USA). Routinely, six perioperative microbiological culture samples were taken from the femoral and tibial interface during surgery to diagnose possible septic revisions. Infected TKA were revised in two stages. The first stage involved removal of all components and cement, aggressive surgical debridement, pulsed lavage and ultimately placement of a temporary antibiotics loaded spacer. During the second stage (when infection was treated successfully) the revision components were placed. Tibial wedges and femoral augments were used to restore bone loss. All patients received standard postoperative rehabilitation [33].

Measurements

Outcome measure

The functional knee society score (fKSS) was modeled as outcome measure [12]. The KSS was assessed by the orthopaedic surgeon or resident during regular follow-up visits prior to the revision surgery, at 3 months following surgery, and at 1–2–5 years post-TKA revision. A KSS subscale score of < 60 points was characterized as poor, 60–69 points as fair, 70–79 points as good, and ≥ 80 points as excellent [5]. In standard literature, the minimal clinically important difference (MCID) is defined as the minimal change on a score that is important to the patient. For patients after a primary total or unicompartmental knee arthroplasty, the MCID ranged between 6.1 and 34.5 points [8, 13, 19]. Successful treatment thresholds of fKSS ranged between 72.5 and 85.5 1 year after TKA [13]. As the MCID estimates for the fKSS in rTKA has not yet been identified, we defined a beneficial outcome 5 years after rTKA as an improvement of ≥ 20 points in fKSS, which is similar to a two categories increase (from poor to good/excellent outcome) in the fKSS scale, or an absolute threshold of fKSS ≥ 80 points (excellent outcome).

Potential predictive factors

Potential predictive factors included age, gender, femoral and tibial bone defects, reason for revision, comorbidity, and preoperative pre-revision scores such as visual analog scale for pain (VAS pain), range of motion, fKSS and cKSS. Comorbidity was assessed using the American Society of Anesthesiologists score (ASA score) [4]. Severe stiffness as a reason for revision was scored positive in patients with limited range of motion (< 70°), with or without pain [33]. Bone defects evaluated during surgery were scored with the Anderson Orthopaedic Research Institute (AORI) classification system, and dichotomized to type I or IIA versus IIB or III [6].

The study protocol was assessed by the regional medical ethical committee [CMO Arnhem–Nijmegen (no. 2003/173)]. Ethical approval was waived by the medical...
ethical committee on basis of the Dutch medical research involving human subjects act (WMO). The present study has been performed and reported according to the TRIPOD statement for the reporting of multivariable prediction models [3].

Data analysis

Descriptive statistics were used to summarize the data. Data were missing for four of the seven potential predictive variables and the outcome, ranging between 0.3–10% (Table 1). These missing values were imputed using multiple imputation with the chained equations procedure (predictive mean matching) [32]. Missing data were assumed to be missing at random (MAR), and ten imputed datasets were created.

Logistic regression was used to evaluate the association between each predictive factor and the outcome. Potential predictors were entered into a logistic regression model taking into account the multiple imputed datasets. Multivariable logistic regression with a backward selection procedure was applied to achieve the most informative and parsimonious combination of predictors. Akaike’s information criterion ($p < 0.157$) was used as selection criterion [36]. The multivariable odds ratios calculated from the logistic regression analysis served evaluating their individual contribution.

The probability ($P$) of having a beneficial functional outcome was constructed with the following formula:

$$P_{\text{beneficial functional outcome}} = \frac{e^{(\hat{b}_0+\hat{b}_1x_1+\hat{b}_2x_2+\cdots+\hat{b}_nx_n)}}{1+e^{(\hat{b}_0+\hat{b}_1x_1+\hat{b}_2x_2+\cdots+\hat{b}_nx_n)}}$$

| Table 1 Characteristics of patients at baseline for revision TKA and outcome at 5 years postoperative |
|--------------------------------------|----------------|----------------|----------------|
| Characteristic | Total sample ($n=295$) | Patients with successful functional KSS ($n=157$) | Patients with unsuccessful functional KSS ($n=119$) |
| n (%) | Missings n (%) | n (%) | Missings n (%) | n (%) | Missings n (%) |
|---|---|---|---|---|---|
| Gender | 295 | – | 157 | – | 119 | – |
| Female | 198 (67.1) | 95 (60.5) | 87 (73.1) | 97 (32.9) | 62 (39.5) | 32 (26.9) |
| Male | 97 (32.9) | 62 (39.5) | 32 (26.9) |
| Age (years)$^a$ | 65 (36–84) | 1 (0.3) | 63.0 (36–84) | – | 67.0 (42–83) | 1 (0.8) |
| KSS subscale$^a$ | | | | | |
| Function | 50 (–20 to 90) | 11 (3.7) | 40.0 (–20 to 90) | 3 (1.9) | 50.0 (–10 to 90) | – |
| Clinical | 52 (–14 to 95) | 11 (3.7) | 52.0 (–8 to 90) | 5 (3.2) | 52.0 (–14 to 95) | 2 (1.7) |
| VAS pain score$^a$ | 64 (0–100) | 20 (6.8) | 63.5 (0–100) | 9 (5.7) | 65.5 (0–100) | 7 (5.9) |
| Reason for revision | 295 | – | 157 | – | 119 | – |
| Aseptic loosening | 76 (25.8) | 42 (26.8) | 30 (25.2) |
| Infection | 60 (20.3) | 35 (22.3) | 20 (16.8) |
| Instability | 52 (17.6) | 27 (17.2) | 22 (18.5) |
| Malposition | 77 (26.1) | 40 (25.5) | 32 (26.9) |
| Severe stiffness | 30 (10.2) | 13 (8.3) | 15 (12.6) |
| Femoral bone defect | 270 | 25 (8.4) | 141 | 16 (10.2) | 111 | 8 (6.7) |
| Type I/IIa | 211 (78.1) | 117 (83.0) | 81 (73.0) |
| Type IIb/III | 59 (21.9) | 24 (17.0) | 30 (27.0) |
| Tibial bone defect, type IIb/III | 266 | 29 (9.8) | 137 | 20 (12.7) | 111 | 8 (6.7) |
| Type I/IIa | 240 (90.2) | 124 (90.5) | 102 (91.9) |
| Type Ib/IIb | 26 (9.8) | 13 (9.5) | 9 (8.1) |
| ASA | 275 | 20 (6.8) | 146 | 11 (7.0) | 112 | 7 (5.9) |
| ASA I | 126 (45.8) | 72 (49.3) | 46 (41.1) |
| ASA II | 93 (33.8) | 46 (31.5) | 41 (36.6) |
| ≥ ASA III | 56 (20.4) | 28 (19.2) | 25 (22.3) |

Outcomes 5 years postoperative

KSS subscale$^a$

| Characteristic | Total sample ($n=295$) | Patients with successful functional KSS ($n=157$) | Patients with unsuccessful functional KSS ($n=119$) |
| n (%) | Missings n (%) | n (%) | Missings n (%) | n (%) | Missings n (%) |
|---|---|---|---|---|---|
| Function | 60 (–20–100) | 11 (3.7) | 80.0 (10–100) | – | 50.0 (–20–75) | – |
| Clinical | 80.5 (5–100) | 29 (9.8) | 90.0 (5–100) | 13 (8.3) | 67.0 (30–100) | 8 (6.7) |

$^a$Data presented as median (range)

$^b$Successful functional outcome is defined as a fKSS ≥ 80 or an increase in fKSS ≥ 20 5 years postoperative. There were 19 patients with a missing outcome.
where $\beta_0$ represents the constant, and $\beta_1$, $\beta_2$ and $\beta_n$ the regression coefficients of the predictors $x_1$, $x_2$, and $x_n$, respectively.

Model performance was assessed with calibration plots. The model’s ability to discriminate between patients with successful and unsuccessful outcomes was estimated by the area under the curve (AUC) of the receiver operating characteristics (ROC) curve of the model. Prediction models derived with multivariable regression analyses are known for overfitting, which results in too extreme predictions when applied in new cases. Therefore, the model was validated internally using bootstrapping techniques for which 500 samples were drawn with replacement from the development sample. Bootstrapping techniques provide information on the performance of the model in comparable datasets and generate a shrinkage factor to adjust the regression coefficients and the intercept. After this adjustment, the model performance was reevaluated. No formal sample size calculation was performed. When multivariable prediction models are developed, the sample size is often based on the ratio of the number of individuals with the outcome event to the number of candidate predictors (more precisely, the number of parameters), referred to as the events per variable (EPV). On the basis of some empirical investigations a rule of thumb to have at least 10 EPV was suggested that has been widely adopted [3]. According to this rule of thumb, we could consider a maximum of 11 variables (119 outcome events/10 EPV).

A nomogram was created to easily calculate the probability of having a beneficial functional outcome 5 years after rTKA for a given patient. Statistical analyses were performed using SPSS [version 22 (IBM, New Jersey, US)] and R [version 3.4.3 (R Foundation, Vienna, Austria)] with package ‘rms’.

**Results**

The 295 analyzed patients had a median age of 65 years (range 36–84), and were mainly female (198; 67.1%). The median preoperative fKSS score was 50 (range 20 to 90); the median 5-year postoperative score was 60 (range 20 to 100). A beneficial fKSS outcome 5 years after rTKA ($\geq 20$ fKSS improvement or absolute fKSS $\geq 80$ values) was observed in 157 patients (53%). The preoperative and postoperative KSS scores, ROM, VAS and the other baseline characteristics are presented in Table 1.

**Predictive ability**

After application of the backward selection procedure, the following variables remained in the multivariable prediction model: age at surgery, gender, femoral bone defects, preoperative fKSS, and severe stiffness as reason for revision (Table 2).

The area under the ROC curve determining how well the model distinguished between patients with beneficial and non-beneficial fKSS 5 years after rTKA was 0.76, 95% CI 0.70–0.81. Through bootstrapping a shrinkage factor of 0.84 was found. After multiplying the regression coefficients with this factor and subsequently adjusting the intercept, the model’s performance was reevaluated. The AUC after internal validation for fKSS was 0.76, 95% CI 0.70–0.81, indicating a good discrimination. Graphical assessment of the model’s calibration after internal validation is shown in Fig. 1.

Male patients had a higher chance of a successful functional outcome 5 years after rTKA surgery than female patients (OR 1.59, 95% CI 0.91–2.78). Patients with bone defects AORI Type IIb/III (OR 0.44, 95% CI 0.22–0.85) patients of higher age (interquartile range OR 0.39, 95% CI 0.26–0.58), patients with higher preoperative fKSS (interquartile range OR 0.42, 95% CI 0.30–0.59), and patients with severe stiffness as reason for revision (OR 0.48, 95% CI 0.20–1.15) had a lower chance of a successful outcome (Table 2).

The probability of having a beneficial functional outcome can be calculated as follows using the regression coefficients presented in Table 2:

$$P_{\text{beneficial functional outcome}} = \frac{e^{lp}}{1 + e^{lp}}$$

where $lp = 6.90 + (-0.07 \times \text{age at surgery}) + (-0.44 \times \text{female}) + (-0.03 \times \text{preoperative fKSS}) +$.

| Predictors                  | Functional KSS | Regression coefficient ($\beta$) after internal validation | Odds ratio (95% CI) |
|-----------------------------|----------------|-------------------------------------------------------------|---------------------|
| Intercept                   | 6.90           | 0.39 (0.26–0.58)                                            | REF                 |
| Age at surgery*             | $-0.07$        | 0.91 (0.63–1.30)                                            | REF                 |
| Gender                      | Male REF       | REF                                                          | REF                 |
|                             | Female $-0.44$ | 0.91 (0.63–1.30)                                            | REF                 |
| Preoperative fKSS*          | $-0.03$        | 0.91 (0.63–1.30)                                            | REF                 |
| Severe stiffness as reason for revision | No REF | REF | REF |
|                             | Yes $-0.68$    | 0.91 (0.63–1.30)                                            | REF                 |
| Femoral bone defect         | Type I/IIa REF | REF                                                          | REF                 |
|                             | Type IIb/III $-0.77$ | 0.91 (0.63–1.30)                                           | REF                 |

*Odds ratios (OR) for continuous predictors are presented as interquartile range OR.
(−0.68 × severe stiffness) + (−0.77 × Type IIb/III femoral bone defect).

The nomogram created as a tool to easily calculate the probability of having a beneficial functional outcome 5 years after rTKA for a given patient is shown in Fig. 2.

**Clinical example**

To illustrate the prediction rule we will predict the outcome for a 65-year-old female patient with a preoperative fKSS of 50 who has a type IIb AORI femur bone defect, and severe stiffness as the main reason for revision. We can calculate the predicted probability of a beneficial fKSS after 5 years after rTKA calculated with the formula presented above or using the nomogram (Fig. 2). The predicted probability for this patient is 26% (95% CI 17–35%), i.e., a small chance of a successful fKSS outcome after 5 years. With this information, the patient is provided with more realistic expectations regarding the revision surgery and its late-anticipated poor results. An otherwise good candidate would be a healthy 55-year-old man with preoperative fKSS of 50 points, type IIa AORI femur bone defect, and a different reason than severe stiffness (e.g. infection) for his revision surgery. His probability of a successful fKSS 5 years after TKA revision is 82% (95% CI 75–90%).

**Discussion**

The most important finding of this study was that the functional outcome 5 years after rTKA could be adequately predicted using preoperative fKSS, age, gender, stiffness as a reason for revision and the type of femoral bone defect. Combining these factors into a relevant prognostic profile will enable physicians to help manage patients’ expectations about revision surgery.

Preoperative fKSS was inversely related to successful 5-year postoperative fKSS outcome, which was earlier described for a 2-year follow-up period [33], and women had a worse outcome than men after a rTKA. This might be because women have higher rates of obesity even morbid obesity, postoperative transfusion, and extended length of stay, all of which are correlated with poor functional and clinical outcomes [34]. However, in a small retrospective study, no differences in gender were found [27]. The coexistence of major femoral bone effects was associated with worse fKSS outcome. This is in line with a study that used bone allograft to fill the bone defects, whereby 10% of all revision required a re-revision [7]. Unfortunately, bone defects can only be adequately scored perioperatively. Nevertheless, patients’ expectations could adequately be managed preoperatively by providing fKSS estimations with or without major femoral bone defects. In our study, the majority of patients had minor femoral bone defects.

While the minimum clinically important difference for revision surgery is unknown, a definition of beneficial outcome of ≥ 20 for fKSS was used. This definition is in the range of MCIDs reported for primary TKA [8, 13, 19]. In addition, ≥ 80 fKSS was set as a clear threshold to ensure patients who already have a high preoperative KSS score to be scored as a successful treatment [13]. Despite this threshold, most benefits was achieved in patients with poor preoperative fKSS. This is in contrast to a previous smaller study, which did not include a proper MCID and did not adjust for
severe stiffness as a reason for revision [14]. Patients with severe stiffness had a poor fKSS outcome in our cohort. Infection was not found to be a predictor for inferior outcomes, which might be surprising. It has been shown that infection as the reason for revision was associated with an inferior functional outcome 2 years after surgery [33]. However, Meek et al. [20] reported that revision operation for infection was associated with reasonable function and satisfaction scores at a mean follow-up of 41 months (minimal follow-up of 2 years). In addition, Patil et al. [25] compared the clinical outcomes and patient satisfaction rates of aseptic versus septic revision TKA at a mean follow-up of 40 months (range 24–80 months). They reported that patients undergoing revision for an infected TKA had better functional outcomes compared to those with revision for aseptic reasons. Revision for severe stiffness was associated with the poorest outcomes.

The risk calculation based on the prediction model can aid the shared decision making when considering a revision TKA. Functional and clinical outcomes from the KSS system are both of interest to orthopedic surgeons, although clinical KSS can be difficult to interpret for patients. The fKSS can be understood as an improvement or decline of their walking ability, including climbing stairs and the use of walking aids. The prediction rule provides a probability between 0% and 100% for beneficial KSS outcome 5 years after revision total knee arthroplasty. However, a 50% probability of beneficial outcome might still leave the patient in doubt as to what to expect. It is important to correctly set the indication for revision. For example, if a TKA is infected,
this alone would justify a revision, even if the outcome prediction is poor. However, the model can still be used to guide patients’ expectations in these situations.

This study has some potential limitations that need to be discussed. The study was conducted in a relatively young population, with few comorbidities. Differences in case-mix, for instance, may result in introducing new predictors that in turn could further improve the performance of the model and resulting in a better application in daily practice and for general use. To make the model easier to use in daily practice, only easily accessible predictors were included. Furthermore, it would have been preferable to consider other potential predictors such as body mass index, diabetic status, smoking, mental status, and time between primary and revision TKA to be included in the model [21, 24, 29]. However, these variables were not recorded. Evidently, the outcome can also be affected by factors as perioperative findings or complications during surgery, or injuries after surgery. These unknown factors were also not included in our model.

Missing values were imputed using multiple imputation techniques, which further improved the predictive ability of the prognostic model [11, 31]. Sample size was limited, which introduced overfitting. Internal validation with bootstrapping techniques were used to adjust for this overfitting [11, 31].

Patients were evaluated 5 years after surgery. Truncating the follow-up data at 2 years would have increased the number of patients, but at the cost of missing relevant changes in KSS. In addition, a longer follow-up gives patients more relevant expectations regarding the clinical process [23]. Before integrating this prediction rule in daily practice, the prediction rule needs to be externally validated, and important determinants such as BMI, smoking and diabetes mellitus should be taken into consideration as well. In daily clinical practice, prediction of functional outcome after rTKA based on the presented prediction model may help clinicians to enable better management of patient expectations and aid shared decision making when considering a rTKA.

**Conclusion**

Future functional outcome among patients who need a revision of their total knee arthroplasty can be predicted using easily accessible patient characteristics. Patients are more likely to achieve a beneficial outcome if they are male, younger, have lower preoperative KSS, and reasons for revision other than severe stiffness. Combining these factors into a relevant prognostic profile will enable physicians to help manage patients’ expectations about revision surgery.

**Acknowledgements** We would like to show our gratitude to the patients who made this study possible.

**Author contributions** All authors (JFMV, GH, KCD, ABW, PJCH) designed the study and interpreted data. JFMV and GH made the statistical analyses. JFMV, GH, and PJCH wrote the first version of the manuscript. All authors critically revised different versions of the manuscript, and all authors read the final version of the manuscript. GH is the corresponding author.

**Funding** The Institute (with authors JFMV, KCD, ABW, and PJCH) received funding from Smith & Nephew for staff to start the prospective database. Smith & Nephew had no involvement in the present study design, data collection, analysis and interpretation, writing of the manuscript and the decision to submit the work for publication.

**Data availability** The datasets generated and/or analysed during the current study are not publicly available but are available from the corresponding author (GH) on reasonable request.

**Compliance with ethical standards**

**Conflict of interest** The authors declare that they have no conflict of interest.

**Ethical approval** The study protocol was assessed by the regional Medical Ethical Committee (CMO Arnhem-Nijmegen (no. 2003/173)). Ethical approval was waived by the Medical Ethical Committee on basis of the Dutch Medical Research Involving Human Subjects Act (WMO).

**Informed consent** Written informed consent was obtained from all study participants.

**Consent to publish** Written consent for publication of the results was obtained from all study participants.

**OpenAccess** This article is distributed under the terms of the Creative Commons Attribution 4.0 International License (http://creativecommons.org/licenses/by/4.0/), which permits unrestricted use, distribution, and reproduction in any medium, provided you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons license, and indicate if changes were made.

**References**

1. Abdel MP, Della Valle CJ (2016) The surgical approach for revision total knee arthroplasty. Bone Jt J 98-B:113–115
2. Baier C, Luring C, Schaumburger J, Kock F, Beckmann J, Tingart M, Zeman F, Gritka J, Springorum HR (2013) Assessing patient-oriented results after revision total knee arthroplasty. J Orthop Sci 18:955–961
3. Collins GS, Reitsma JB, Altman DG, Moons KG (2015) Transparent reporting of a multivariable prediction model for individual prognosis or diagnosis (TRIPOD): the TRIPOD statement. BMJ 350:g7594
4. Davenport DL, Bowe EA, Henderson WG, Khuri SF, Mentzer RM Jr (2006) National surgical quality improvement program (NSQIP) risk factors can be used to validate American Society of Anesthesiologists Physical Status Classification (ASAPS) levels. Ann Surg 243:636–641
5. Dowsey MM, Choong PF (2013) The utility of outcome measures in total knee replacement surgery. Int J Rheumatol 2013:506518
6. Engh GA, Ammeen DJ (1998) Classification and preoperative radiographic evaluation: knee. Orthop Clin North Am 29:205–217
7. Franke KF, Nusem I, Gamboa G, Morgan DA (2013) Outcome of revision total knee arthroplasty with bone allograft in 30 cases. Acta Orthop Belg 79:427–434
8. Giesinger JM, Hamilton DF, Jost B, Behrend H, Giesinger K (2015) WOMAC, EQ-SD and knee society score thresholds for treatment success after total knee arthroplasty. J Arthroplasty 30:2154–2158
9. Giesinger JM, Loth FL, MacDonald DJ, Giesinger K, Patton JT, Simpson A, Howie CR, Hamilton DF (2018) Patient-reported outcome metrics following total knee arthroplasty are influenced differently by patients’ body mass index. Knee Surg Sports Traumatol Arthrosc 26:3257–3264
10. Hardeman F, Londers J, Favril A, Witvrouw E, Bellemans J, Victor J (2012) Predicting factors which are relevant for the clinical outcome after revision total knee arthroplasty. Knee Surg Sports Traumatol Arthrosc 20:1049–1056
11. Harrell FE Jr (2015) Multivariable modeling strategies. In: Regression modeling strategies. Applications to linear models, logistic regression, and survival analysis. Springer, Cham, pp 63–102
12. Insall JN, Dorr LD, Scott RD, Scott WN (1989) Rationale of the knee society clinical rating system. Clin Orthop Relat Res 248:13–14
13. Jacobs CA, Christensen CP (2009) Correlations between knee society function scores and functional force measures. Clin Orthop Relat Res 467:2414–2419
14. Kasmire KE, Rasouli MR, Mortazavi SM, Sharkey PF, Parvizi J (2014) Predictors of functional outcome after revision total knee arthroplasty following aseptic failure. Knee 21:264–267
15. Kurtz S, Kang O, Lau E, Mowat F, Halpern M (2007) Projections of primary and revision hip and knee arthroplasty in the United States from 2005 to 2030. J Bone Jt Surg Am 89:780–785
16. Kurtz SM, Lau E, Ong K, Zhao K, Kelly M, Bozic KJ (2009) Future young patient demand for primary and revision joint replacement: national projections from 2010 to 2030. Clin Orthop Relat Res 467:2606–2612
17. Kurtz SM, Ong KL, Lau E, Bozic KJ (2014) Impact of the economic downturn on total joint replacement demand in the United States: updated projections to 2021. J Bone Jt Surg Am 96:624–630
18. Labek G, Thaler M, Janda W, Agreiter M, Stockl B (2011) Revision rates after total joint replacement: cumulative results from worldwide joint register datasets. J Bone Jt Surg Br 93:293–297
19. Lee WC, Kwan YH, Chong HC, Yeo SJ (2017) The minimal clinical important difference for knee society clinical rating system after total knee arthroplasty for primary osteoarthritis. Knee Surg Sports Traumatol Arthrosc 25:3354–3359
20. Meek RM, Masri BA, Dunlop D, Garbuz DS, Greidanus NV, McGraw R, Duncan CP (2003) Patient satisfaction and functional status after treatment of infection at the site of a total knee arthroplasty with use of the PROSTALAC articulating spacer. J Bone Jt Surg Am 85A:1888–1892
21. Mulhall KJ, Ghomravi HM, Mihalko W, Cui J, Saleh KJ (2007) Adverse effects of increased body mass index and weight on survivorship of total knee arthroplasty and subsequent outcomes of revision TKA. J Knee Surg 20:199–204
22. Namba RS, Cafri G, Katod M, Inacio MC, Brox TW, Paxton EW (2013) Risk factors for total knee arthroplasty aseptic revision. J Arthroplasty 28:122–127
23. Nikolaus OB, McLendon PB, Hансsen AD, Mabry TM, Berbari EF, Sierra RJ (2016) Factors associated with 20-year cumulative risk of infection after aseptic index revision total knee arthroplasty. J Arthroplasty 31:872–877
24. Nwachukwu BU, Gurary EB, Lerner V, Collins JE, Thornhill TS, Losina E, Katz JN (2015) Effect of smoking and soft tissue release on risk of revision after total knee arthroplasty: a case-control study. BMC Musculoskelet Disord 16:245
25. Patil N, Lee K, Huddleston JL, Harris AH, Goodman SB (2010) Aseptic versus septic revision total knee arthroplasty: patient satisfaction, outcome and quality of life improvement. Knee 17:200–203
26. Paxton EW, Inacio MC, Khatod M, Yue E, Funahashi T, Barber T (2015) Risk calculators predict failures of knee and hip arthroplasties: findings from a large health maintenance organization. Clin Orthop Relat Res 473:3965–3973
27. Pun SY, Ries MD (2008) Effect of gender and preoperative diagnosis on results of revision total knee arthroplasty. Clin Orthop Relat Res 466:2701–2705
28. Sheng P, Lehto M, Kataja M, Halonen P, Moilanen T, Pajamaki J (2004) Patient outcome following revision total knee arthroplasty: a meta-analysis. Int Orthop 28:78–81
29. Singh JA, Lewallen DG (2013) Diabetes: a risk factor for poor functional outcome after total knee arthroplasty. PLoS One 8:e78991
30. Springer BD, Hanssen AD, Sim FH, Lewallen DG (2001) The kinematic rotating hinge prosthesis for complex knee arthroplasty. Clin Orthop Relat Res 392:283–291
31. Steyerberg EW (2009) Overfitting and Optimism in Prediction Models. In: Clinical prediction models. A practical approach to development, validation, and updating. Springer Science + Business Media LLC, New York, pp 83–99
32. van Buuren S (2012) Multivariate missing data. In: Flexible imputation of missing data. Chapman and Hall/CRC, Boca Raton, pp 95–121
33. van Kempen RW, Schimmel JJ, van Hellemontd GG, Vanden-neucker H, Wymenua AB (2013) Reason for revision TKA predicts clinical outcome: prospective evaluation of 150 consecutive patients with 2-years followup. Clin Orthop Relat Res 471:2296–2302
34. Whitlock KG, Piconov HI, Shah SH, Wang OJ, Gonzalez MH (2016) Gender role in total knee arthroplasty: a retrospective analysis of perioperative outcomes in US patients. J Arthroplast 31:2736–2740
35. Wilke BK, Wagner ER, Trousdale RT (2014) Long-term survival of semi-constrained total knee arthroplasty for revision surgery. J Arthroplast 29:1005–1008
36. Wood AM, White IR, Royston P (2008) How should variable selection be performed with multiply imputed data? Stat Med 27:3227–3246