Measuring Online Social Support: Development and Validation of a Short Form for Chinese Adolescents

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Abstract: Supportive interactions on social media have great potential to benefit adolescents’ development. However, there is no instrument to measure online social support (OSS) in China. The study aimed to develop and validate a Chinese short version of the Online Social Support Scale (OSSS). The original scale was translated into Chinese through multiple forward and backward translation protocols. The calibration sample (N = 262) was used to select items and test the reliability, validity, and internal structure of the short form. The cross-validation sample (N = 267) was then used to assess measurement invariance by multigroup confirmatory factor analysis and examine criterion validity based on its relationships with life satisfaction, depression, and time on social media. The 20-item Chinese short version of OSSS (OSSS-CS) includes four factors: esteem/emotional support, social companionship, informational support, and instrumental support. Our results suggest that the OSSS-CS has high internal consistency, construct validity, and criterion validity. Furthermore, evidence of partial cross-validity demonstrated invariance of the variance–covariance matrices, factor structure, factor loadings, and factor variance across independent samples. The results also revealed that the original OSSS could be replicated across cultures. Finally, the short form developed in the study can be used as a reliable and valid measure of online social support among the Chinese adolescent population.

Keywords: adolescent; social support; online social support; confirmatory factor analysis; scale validation; cross-validation

1. Introduction

In the digital era, the Internet plays an enormous role in people’s daily lives. Social media platforms have become a global means of social interaction for people of all ages, especially among the younger populations. The use of social media in China is particularly ubiquitous and popular [1]. In 2021, China became the world’s largest social media market, attracting 926 million active social media users [2]. Among them, Chinese adolescents account for a relatively high proportion [3]. Social media is the second-most commonly used smartphone application among Chinese adolescents, with an 88.7% share, just after music apps at 95.8% [4]. It is worth noting that popular international social media such as Facebook, Instagram, and Twitter are blocked in China. Chinese users gather on a completely different set of social media channels such as WeChat, QQ, and Weibo [5]. The independent social media ecosystem in China underscores how different it is from its Western counterparts and the value and importance of studying social interactions in China’s mediascape.

With the proliferation of social media, the way in which adolescents interact has changed dramatically. Scholars have taken note of the trend, resulting in a growing body of research on this topic. While some studies have revealed the potential dangers and adverse impacts of social media, such as cyberbullying and Internet addiction [6–8], far less research has focused on the positive aspects of social media [9]. With the pervasive use of social media, it has come to represent new and important sources of or vehicles...
for social support [10]. Social media may fill a need for social belongingness, provide a
distraction from various stressors, or enable people to express and receive support via
common features, such as being “friended” or “followed” by others, liking a post, or
leaving an encouraging comment [11,12]. A Chinese survey involving 5076 adolescents
in six provinces showed that more than half (59%) of the adolescents agreed with the
statement, “The Internet brings people closer,” compared with 38.3% of adolescents who
agreed with the statement, “The Internet makes people more alienated” [13]. Despite the
fact that online social relations convey opportunities and risks, research on their positive
side is still in the early stages [9].

Social support is fundamental nourishment in a person’s environment and is crucial to
many development stages [14]. The positive role of social support on mental and physical
health has been widely acknowledged [15]. In childhood and adolescence, social support is
a contributing factor to positive development and general health behaviors, a protective
factor from psychological distress, and a predictor of well-being in adulthood [16]. In
the traditional face-to-face context, the main effect model suggests that social resources
have a beneficial effect on mental health, irrespective of whether or not persons are under
stress [15]. The social support generated from online space, so-called online social support
(OSS), can be defined as the Internet-facilitated receipt of emotional, informational and
instrumental assistance, and sense of belonging from friends, family, and others in one’s
social networks [17,18]. OSS appears to have promising influence on adolescents’ well-
being and operates in ways that are quite similar to in-person social interactions [19]. A
systematic review found that Facebook-based social support improved well-being and
general physical and mental health. It was also found to reduce mental illness, including
depression, anxiety, and loneliness [20]. However, some evidence showed that social
support from social media was related to more significant mental health conditions, e.g., a
higher likelihood of depression and poorer quality of life [12,21]. The inconsistent findings
regarding the relationship between OSS and adolescents’ mental health were partly because
of the varied conceptualization and operationalization of OSS [22].

Current measurements of OSS are mainly altered versions of established face-to-face
social support measures based on different theoretical conceptualizations of social support.
For example, some studies measured perceptions of online support, i.e., the personal
interpretation of supportive transactions, such as the Interpersonal Support Evaluation
List [23] used by Indian and Grieve [24] and the family subscale of the Multidimensional
Scale of Perceived Social Support [25] used by Frison and Eggermont [26]. Some studies
measured received social support, i.e., the amount of social support experienced, such
as the Inventory of Socially Supportive Behaviors [27] used by Kim [28] and the tone
of reactions in Valkenburg et al. [29]’s study. Additionally, Lin et al. [6] used the Social
Support Scale [30] to capture OSS from people acquainted only through the Internet, as
distinguished from people known in person. There are several newly developed scales
specifically for OSS, such as the Online Social Support Scale (OSSS) [9], which is used
to measure the frequency of support behaviors, and the Liang OSS Scale [31], which
measures social support perceptions. However, these extant measures tend to rely on ad
hoc measures, often without strong psychometric evidence, and their varied theoretical
foundation has complicated interpretations [9].

Among these scales, the OSSS is the first to have both theoretical roots and strong
psychometric support. To capture the multifaceted nature of OSS, Nick and colleagues
conducted a thorough literature review of theoretical and empirical attempts to characterize
the structure of in-person social support [9]. They found four factors that have consistently
arisen in the literature: (1) esteem/emotional support (EE) refers to expressions of accep-
tance, intimacy, care, like, respect, etc. through verbal and nonverbal cues from others;
(2) social companionship (SC) support fosters a sense of belonging through expressions
of inclusivity or spending time together; (3) informational support (INF) includes giving
advice and feedback, sharing information, perspectives, and knowledge, and providing re-
sources; and (4) instrumental support (INS) means providing financial, material assistance,
helping with tasks, and taking responsibility [9]. This classification divided social support according to its function.

Evidence of reliability, convergent validity, and discriminant validity supports the psychometric properties of the OSSS [9]. Its validity and reliability have been proven in college and community samples in the US [9,32]. However, little is known about the psychometric properties of the OSSS in the adolescent population and Eastern cultures. It is possible that the factor structure of the OSSS might vary across ages, as teens can have different social interaction patterns compared with adults. Junior high school students, in particular, are at the early stage of a transitional period and are just beginning to develop their Internet usage patterns and habits. Furthermore, the validation of measures across cultures and languages is an important step in studying human behaviors that are deemed universal. Studying whether this scale is applicable to China’s unique social media ecological system is necessary and helpful to compare the similarities and differences in OSS across societies, cultures, and media platforms.

As a reliable measurement of OSS is relatively scarce in contemporary China, adapting the OSSS to the Chinese context can be a timely effort to facilitate further research on online social relations among Chinese youth. In addition, developing a Chinese version of the OSSS can have significance for further cross-cultural research on adolescent development in the digital era. Although OSSS has high psychometric qualities and has the potential to be broadly used in future studies, one limitation is that it has 40 items. Completing the full scale takes a long time, and this questionnaire may not be practical in a busy setting. Therefore, a short form is needed to make it more feasible and convenient to administer in community and school settings.

In this study, we aimed to provide a validated Chinese short form of the OSSS (OSSS-CS) for use among the adolescent population. First, we translated and adapted the OSSS into Chinese; then we selected items to develop a short version of the OSSS and tested its psychometric properties in the Chinese adolescent population. To rule out the possibility that the factor structure of the OSSS-CS was achieved only through the post hoc selection of statements that met our selection criteria, we administered the OSSS-CS to another sample to test the replicability of the modified model. Finally, previous research suggested that social media use contributed to OSS [9,33], and OSS was correlated with life satisfaction and depression [10,20,26]. Thus, we demonstrated its criterion validity by examining its relationships with social media use, life satisfaction, and depression.

To accomplish these purposes, we considered five hypotheses, as follows:

**Hypothesis 1 (internal structure).** The indicators used to measure OSS reveal four factors that fit the data;

**Hypothesis 2 (reliability).** The scale and subscales show high internal consistency;

**Hypothesis 3 (construct validity).** The indicators of each latent variable have a high shared variance; each latent variable is unique compared with other constructs;

**Hypothesis 4 (measurement invariance).** The scale shows robust invariance across samples;

**Hypothesis 5 (criterion validity).** The dimensions of the OSSS-CS have positive relations with social media use and life satisfaction, and a negative relation with depression.

**2. Materials and Methods**

**2.1. Participants**

The data were collected by a paper-based questionnaire survey conducted between September 2021 and December 2021. Approval to conduct the studies was obtained from the Survey and Behavioral Research Ethics Committee of the Chinese University of Hong Kong. The participants consisted of two groups. The calibration sample comprised 262 students...
in their second year of junior high school from one school in Foshan, Guangdong Province. Students in six classes were invited to participate in the study. A total of 312 students participated, with 262 valid responses. Their ages ranged from 12 to 16 years (mean age = 13, SD = 0.56), and 47.7% were female. The cross-validation sample consisted of 267 students in two other junior high schools in Foshan. One class was randomly chosen from each grade of each school. Thus, a total of six classes were involved, and all students in the chosen classes were invited to participate in the survey. A total of 320 students filled out the questionnaire, with 267 valid responses. This sample comprised 137 (51.3%) males and 130 (48.7%) females, whose ages ranged from 12 to 15 years (mean age 13, SD = 1.00). The demographics of the two separate samples are shown in Table 1. A total of 529 middle school students were included in the study. As can be seen, there are more boys than girls among the participants. The average age of the participants was 13 years old, and over half of the students lived on campus. Around 80% of the students’ household registrations were in Foshan, while the rest were from other regions. The majority of students rated their family income as similar to or better than average. In terms of social media use, students spent more time on social media during weekends and holidays than on weekdays.

Table 1. General characteristics of participants.

|                          | N   | %    |
|--------------------------|-----|------|
| Gender                   |     |      |
| Male                     | 137 | 52.3 |
| Female                   | 125 | 47.7 |
| Grade                    |     |      |
| 1                        | 99  | 37.1 |
| 2                        | 262 | 100  |
| 3                        | 88  | 33.0 |
| Living arrangement       |     |      |
| Residential              | 155 | 59.2 |
| Commuter                 | 107 | 40.8 |
| Household registration   |     |      |
| Local                    | 216 | 82.4 |
| Non-local                | 46  | 17.5 |
| Family income            |     |      |
| High                     | 146 | 55.7 |
| Middle                   | 101 | 38.5 |
| Low                      | 15  | 5.8  |

| Daily time spent on social media | M (hour) | SD |
|---------------------------------|----------|----|
| Weekdays                        | 0.17     | 0.25 |
| Weekends and holiday            | 1.78     | 1.56 |

Note: Calibration sample in regular font, cross-validation sample in italics.

2.2. Measures

The OSSS is a five-point Likert scale comprising 40 items. It has four subscales: esteem/emotional support (e.g., “People encourage me when I’m online”), social companionship (e.g., “When I’m online, I talk or do things with other people”), informational support (e.g., “People help me learn new things when I’m online”), and instrumental support (e.g., “When I’m online, people help me with school or work”). Items inquire about how often the supportive interactions have happened while the participants interacted with others online over the last two months. Each item was scored using a five-point Likert scale, wherein a higher score indicates greater OSS. The scale is well-validated and has a high degree of internal consistency; the coefficient alpha ranged from 0.94 to 0.95 [9].

Depression was measured using the short form of the Center for Epidemiological Studies Depression Scale (CES-D) [34]. The 10-item CES-D demonstrated good reliability and construct validity in the adolescent population [35]. Among Chinese adolescents (aged 11–18), Cronbach’s α was 0.85 China (r = 0.85), which presents a viable alternative to the full instrument [36]. In the CES-D, the participants were asked to consider their past week while answering questions on a scale of 0–3 (0 = rarely or none of the time (less than 1 day), 1 = some or a little of the time (1–2 days), 2 = occasionally or a moderate amount of time (3–4 days), 3 = most or all of the time (5–7 days)). In this study, Cronbach’s α was 0.86.
Life satisfaction was measured with the Satisfaction with Life Scale (SWLS) [37], a five-item self-report instrument intended to assess satisfaction with one’s life as a whole. The respondents indicate their degree of agreement with each statement of the SWLS using a seven-point Likert scale (7 = “strongly agree,” 1 = “strongly disagree”). The SWLS has been extensively used and translated into more than 25 languages, showing good reliability and validity [38]. The instrument has also been shown to have excellent validity and internal consistency (α = 0.76) in the Chinese adolescent population [39]. In the current study, the reliability of the scale was 0.86.

Social media use was measured by asking participants how many hours per day during weekdays (versus weekends and holidays) they spent using social media (e.g., WeChat, QQ, and Weibo).

We also collected some basic information about the participants by self-report. Students were asked to indicate their household registration (whether from Foshan or other areas) and living arrangements (whether living on campus during weekdays) and to rate their family income (as high, middle, or low).

2.3. Translation Procedure

Multiple forward- and backward-translation protocol was used in translation. Two bilingual Social Work PhD students independently translated the OSSS into Chinese. Chinese was the native language of all translators, and they possessed a high level of English language proficiency. Inconsistency was solved through discussion. A third PhD student who was blinded to the original version retranslated the Chinese version to English, and the back-translated scale was compared with the original scale. Only a minor inconsistency was identified, which was further resolved by a discussion between the third and the previous two translators. The pretesting of the Chinese version was performed first among six social workers in youth service and then by 10 junior high school students. They were invited to review the translated scale and make suggestions wherever necessary. The clarity of the items, the relevance of the content to online situations, the simplicity of the instructions, and the capacity of teen respondents to complete the scale were all considered in the review. The tool was finalized based on these discussions. The final version of the 40-item Chinese OSSS was conducted on the calibration sample.

2.4. Data Analysis

First, descriptive statistics of the OSSS-CS and the four subscales were analyzed. Then, confirmatory factor analysis (CFA) was conducted on the calibration sample with a four-factor structure. To select the items of the original measure that best assessed the construct, we followed the guideline of item selection for the short-form scale development of Marsh et al. [40]. The following criteria were used in this study: (1) reducing the length of the OSSS by 50% (i.e., to retain no more than 20 items), (2) retaining the same content coverage of the four dimensions of the OSSS, (3) retaining at least five items in each subscale, (4) maintaining at least similar alpha coefficients for each subscale, (5) achieving an item-total correlation of at least 0.40, (6) providing a factor structure with acceptable goodness of fit indexes, (7) ensuring that the size of standardized item factor loadings must be CFA > 0.4, and (8) checking whether items have minimum cross-loading and correlated uniqueness.

Once the items of the OSSS-CS were selected, we conducted CFA to test the structural validity of OSSS-CS. The goodness of fit indices include a non-normed fit index (NNFI), a comparative fit index (CFI), a standardized root mean square residual (SRMR), and a root mean square error of approximation (RMSEA). The NNFI and CFI values, ≥0.90 and ≥0.95, respectively, were indicative of acceptable and good model fit, as were the corresponding SRMR (≤0.10 and ≤0.08) and RMSEA values (≤0.08 and ≤0.06) [41,42]. We then calculated Cronbach’s α to examine the scales’ internal consistency and computed the correlations between the original measure and the shortened version.

Next, to assess the generalizability of the OSSS-CS, we cross-validated the scale using a multistep procedure. We tested loose, partial, and tight cross-validation based on the
progression of invariance constraints specified by MacCallum et al. [43]. In loose cross-validation, all the parameters in the model are re-estimated in an independent sample (i.e., the cross-validation sample) [44]. In partial cross-validation, a series of nested models were performed with added constraints on item loadings, factor covariances, and uniqueness values successively. The criteria for tight cross-validation are satisfied when all of the parameters in the model are held equal between samples and are able to reproduce the observed covariances in the second sample [43]. The nonsignificant results of the chi-square difference test ($p > 0.01$) [45], $|\Delta CFI| < 0.01$ [46], and $|\Delta RMSEA| < 0.015$ [47] would support the invariance assumption. Finally, Pearson correlations of the total scale and subscale scores with mental health variables and social media use were assessed to verify criterion validity in the cross-validation sample. Data analyses were performed with SPSS 25.0 (IBM: New York, NY, USA) and Amos 23.0 (IBM: Chicago, IL, USA).

3. Results

3.1. Descriptive Analyses

Table 2 presents the means and standard deviations of the variables of interest for both the calibration and cross-validation samples. The collected responses suggested that adolescents exhibited a moderate level of OSS, with instrumental support at a lower level compared with the other three types of OSS. Values’ skewness and kurtosis indicated normal univariate distribution in full-scale scores and subscales’ scores. However, multivariate normality was violated (Mardia’s statistic > 5). As such, in the following CFA, we employed ML estimation using 2000 bootstrap samples to calculate the Bollen–Stine adjustment for the chi-square $p$ value [48].

Table 2. Descriptive statistics for the OSSS-CS across the entire sample.

|        | M    | SD   | Skewness | Kurtosis |
|--------|------|------|----------|----------|
| OSSS-CS| 2.71 | 2.91 | 0.82     | 1.06     | −0.06    | 0.05   | −0.74 |
| EE     | 2.85 | 2.86 | 0.89     | 1.16     | 0.06     | 0.09   | −0.06 | −0.81 |
| SC     | 2.76 | 3.00 | 0.97     | 1.17     | 0.21     | −0.05  | −0.28 | −0.94 |
| INF    | 2.79 | 2.99 | 0.95     | 1.18     | 0.25     | −0.02  | −0.09 | −0.92 |
| INS    | 2.43 | 2.77 | 0.96     | 1.14     | 0.62     | 0.14   | 0.26  | −0.75 |

Note: Calibration sample in regular font, cross-validation sample in italics. Abbreviations: OSSS-CS, online social support scale-Chinese short version; EE, emotional/esteem support; SC, social companionship support; INF, informational support; INS, instrumental support.

3.2. Calibration Sample Analyses

Using the proposed criteria, 20 items were selected to construct a short form scale of the OSSS-CS (emotional support 1, 2, 3, 4, and 7; companionship support 11, 12, 13, 14, and 20; informational support 22, 24, 25, 26, and 27; and instrumental support 32, 34, 36, 37, 39. See Appendix A, Table A1). The total score of the OSSS-CS was highly and significantly correlated with the full scale ($r = 0.99, p < 0.000$), and the OSSS-CS explained 76.51% of the variance of the full scale.

Internal consistency reliability was assessed by calculating Cronbach’s alpha. Results from the calibration sample indicated that the coefficient alpha value for the 20-item OSSS-CS was 0.963, while those for esteem/emotional support, social companionship, informational support, and instrumental support were 0.917, 0.8, 0.935, and 0.922, respectively, suggesting good internal consistency for the OSSS-CS and for each factor.

Next, CFA was conducted for the short version in the calibration sample. The four-factor model of the OSSS-CS fitted the data well ($\chi^2/df = 2.105, p = 0.001$, Bollen–Stine $p = 0.066$, CFI = 0.961, NFI = 0.928, PGFI = 0.065, RMSEA = 0.065, SRMR = 0.044). All factor loadings were statistically significant and above 0.7 (see Table 3). The composite reliability (CR) indicates that all items of each latent variable constantly measure their corresponding construct. In all cases, CR was higher than 0.90. Thus, the indicators of the four factors are reliable measures of the constructs. Table 4 displays the results of convergent and
discriminate validities. Convergent validity is supported if the average variance extracted (AVE) value is greater than 0.50. In this study, the AVE showed values higher than 0.5 in all cases. Thus, a substantial amount of the variance of the indicators is explained by the construct compared with the error of the measurement. Discriminant validity is evident if the square root of every AVE value belonging to each latent construct is much larger than any correlation among any pair of latent constructs [49]. This criterion is fulfilled in the calibration sample.

| Factors | Indicators | Est. | SE. | p  | Factor Loading | SMC | CR | AVE |
|---------|------------|------|-----|----|----------------|-----|----|-----|
| EE      | Item 1     | 1    |     |    | 0.91           | 0.827 | 0.916 | 0.839 |
|         | Item 2     | 0.948 | 0.039 | *** | 0.916           | 0.839 | 0.916 | 0.86 |
|         | Item 3     | 1.013 | 0.040 | *** | 0.927           | 0.86 | 0.921 | 0.704 |
|         | Item 4     | 0.748 | 0.052 | *** | 0.713           | 0.509 | 0.713 | 0.509 |
|         | Item 7     | 0.763 | 0.056 | *** | 0.691           | 0.478 | 0.691 | 0.478 |
| SC      | Item 11    | 1    |     |    | 0.752           | 0.565 | 0.752 | 0.565 |
|         | Item 12    | 1.214 | 0.079 | *** | 0.903           | 0.815 | 0.903 | 0.815 |
|         | Item 13    | 1.218 | 0.082 | *** | 0.882           | 0.778 | 0.882 | 0.778 |
|         | Item 14    | 1.020 | 0.085 | *** | 0.732           | 0.535 | 0.732 | 0.535 |
|         | Item 20    | 1.036 | 0.086 | *** | 0.733           | 0.537 | 0.733 | 0.537 |
| INF     | Item 22    | 1    |     |    | 0.82            | 0.672 | 0.82 | 0.672 |
|         | Item 24    | 1.108 | 0.060 | *** | 0.907           | 0.823 | 0.907 | 0.823 |
|         | Item 25    | 1.054 | 0.065 | *** | 0.835           | 0.698 | 0.835 | 0.698 |
|         | Item 26    | 1.043 | 0.058 | *** | 0.893           | 0.798 | 0.893 | 0.798 |
|         | Item 27    | 1.035 | 0.062 | *** | 0.849           | 0.721 | 0.849 | 0.721 |
| INS     | Item 32    | 1    |     |    | 0.765           | 0.585 | 0.765 | 0.585 |
|         | Item 34    | 1.128 | 0.076 | *** | 0.848           | 0.719 | 0.848 | 0.719 |
|         | Item 36    | 1.216 | 0.075 | *** | 0.91            | 0.829 | 0.91 | 0.829 |
|         | Item 37    | 1.119 | 0.078 | *** | 0.83            | 0.688 | 0.83 | 0.688 |
|         | Item 39    | 1.111 | 0.075 | *** | 0.844           | 0.713 | 0.844 | 0.713 |

Table 3. Results of CFA among calibration sample.

Abbreviations: Est. = estimation, SE. = standard error, SMC = squared multiple correlations, CR = construct reliability, AVE = average variance extracted; *** p < 0.001; EE, emotional/esteem support; SC, social companionship support; INF, informational support; INS, instrumental support.

Table 4. Results of the convergent and discriminant validity of the OSSS-CS.

| AVE | EE | SC | INF | INS |
|-----|----|----|-----|-----|
| EE  | 0.704 | 0.839 |     |     |
| SC  | 0.646 | 0.664 | 0.804 |     |
| INF | 0.742 | 0.727 | 0.734 | 0.861 |
| INS | 0.707 | 0.662 | 0.715 | 0.831 | 0.841 |

Abbreviations: AVE, average variance extracted; EE, emotional/esteem support; SC, social companionship support; INF, informational support; INS, instrumental support.

3.3. Cross-Validation Sample Analyses

To test the generality of the adapted short version, reliability and construct validity were examined in the cross-validation sample, and the measurement invariance was assessed across the two samples.

The results from the cross-validation sample showed that the coefficient alpha value for the 20-item OSSS-CS was 0.969, while those for esteem/emotional support, social companionship, informational support, and instrumental support were 0.944, 0.900, 0.932, and 0.913, respectively. Therefore, the results from the cross-validation sample indicated the high internal consistency reliability of the short form of the OSSS-CS. In the cross-validation sample, CR was higher than 0.9, and the AVE value was greater than 0.60. The criteria of composite reliability and convergent validity were fulfilled. Furthermore, the criterion of
discriminate validity was achieved in all but two cases. Specifically, the square roots of the AVE of INF (0.859) and INS (0.822) were slightly lower than the correlation between them. Likewise, the square root of the AVE of SC (0.806) was slightly lower than its correlations with INF (0.859) and INS (0.822).

To assess the cross-validity of the short scale adapted from the calibration sample, a series of progressively more restrictive models was tested (Table 5). First, loose cross-validation was performed. This model was specified as described previously and fitted the data well. Then, the equal variance-covariance matrices model was tested on the total sample (combination of calibration and cross-validation samples), and the model fitted well. Factor structure, item loadings, factor variances-covariances, and uniqueness values were sequentially constrained to be equal. The chi-square difference tests between these nested models were statistically significant and sensitive to sample size. Changes in CFI and RMSEA were lower than the thresholds except for uniqueness, suggesting that the factor structure, factor loadings, and factor variances were invariant across the calibration and cross-validation samples.

Table 5. Fit Indexes for cross-validation and invariance analyses.

| Model                                           | df  | $\chi^2$ | RMSEA | CFI   | SRMR |
|-------------------------------------------------|-----|----------|-------|-------|------|
| Cross-validation sample                         | 164 | 402.226  | 0.074 | 0.953 | 0.040|
| Equivalent covariance matrices                  | 164 | 412.617  | 0.054 | 0.974 | 0.033|
| Equivalent factor structure                     | 328 | 747.514  | 0.049 | 0.957 | 0.044|
| Fixed structure and item loading                | 344 | 791.287  | 0.050 | 0.954 | 0.042|
| Fixed structure, item loading, and factor variance/covariance | 354 | 823.343  | 0.050 | 0.951 | 0.077|
| Fixed structure, item loading, factor variance/covariance, and uniqueness | 374 | 1098.811 | 0.061 | 0.925 | 0.059|

Model Comparison

| Item loadings-structure                          | 16  | 43.773   | 0.001 | 0.003 | 0.000|
| Factor variances/covariances-item loadings      | 10  | 32.056   | 0.000 | 0.003 | 0.000|
| Uniqueness-factor variance/covariance           | 20  | 275.468  | 0.011 | 0.026 | 0.000|

Abbreviations: RMSEA, root mean square error of approximation; SRMR, standardized root mean square residual; CFI, comparative fit index.

Criterion validity was estimated by correlating the scores of the OSSS-CS and the four subscales with participants’ mental health (i.e., depression and life satisfaction) and social media use in the cross-validation sample. As shown in Table 6, life satisfaction was positively correlated with the OSSS-CS and the four dimensions, while depression was only negatively related to informational support ($r = -0.124, p < 0.05$) and instrumental support ($r = -0.128, p < 0.05$). Adolescents’ social media use was significantly related to OSS, regardless of whether this occurred on weekdays or weekends/holidays, but the correlation was stronger during the latter.

Table 6. Results of criterion validity among the cross-validation sample.

| OSSS   | EE         | SC         | INF        | INS        |
|--------|------------|------------|------------|------------|
| Depression | -0.111    | -0.087    | -0.065    | -0.124 *   | -0.128 *   |
| Life satisfaction | 0.283 **  | 0.255 **  | 0.226 **  | 0.277 **   | 0.267 **   |
| Social media use | Weekdays  | 0.136 *   | 0.098     | 0.145 *    | 0.119      | 0.132 *    |
|        | Weekend & holiday | 0.173 ** | 0.140 *   | 0.149 *    | 0.153 *    | 0.186 **   |

* $p < 0.05$, ** $p < 0.01$. Abbreviations: OSSS, online social support scale; EE, emotional/esteem support; SC, social companionship support; INF, informational support; INS, instrumental support.
4. Discussion

Social media has become an important part of adolescents’ daily lives, providing a new social context for their interactions [50]. In recent years, the literature has begun to document the positive effects of online social interaction [9]. The accurate measurement of OSS is critical for practitioners and researchers to improve our understanding of OSS and capture its benefits. The OSSS is the first OSS measurement that has both deep theoretical roots and strong psychometric properties [9]. The current study developed a short version of OSSS to be used in the Chinese adolescent population.

The results of the CFA support Hypothesis 1, which states that the OSSS-CS reserved the full scale’s four-factor structure: esteem/emotional support, social companionship, informational support, and instrumental support. The findings demonstrated that the subtypes of social support that apply to the US college and university populations also apply to the Chinese adolescents being studied. The overall OSSS-CS and the four subscales displayed good reliability in the two samples, thus supporting Hypothesis 2. Convergent validity was supported by AVE values, and discriminant validity was confirmed in the calibration sample as the AVE in each factor was greater than the square of this correlation with the other scale factors. However, in the cross-validation sample, the square root of the AVE of SC was slightly lower than its correlation with other latent constructs, suggesting that adolescents may feel companionship from other supportive activities. Thus, it might be difficult to distinguish social companionship support from other dimensions. Future studies may benefit from qualitative studies to reveal the meaning that adolescents attach to OSS.

The measurement invariance hypothesis was partially supported. Although the results demonstrated invariance of the covariance matrices, factor structure, factor loadings, and factor variance across two independent samples, the criteria for tight cross-validation are not satisfied as uniqueness values were not equal between the groups. Tight cross-validation is the strongest evidence for cross-validity, but because the measurement and specific errors inherent in uniqueness values are assumed to be random across samples, it may not be realistic to achieve [43]. Thus, when the item loadings and factor covariances can be constrained to equality across samples, partial cross-validation evidence for a measure is considered quite acceptable.

The criterion validity hypothesis was partly supported. In particular, the results demonstrated a positive correlation between OSSS-CS and life satisfaction. This is in line with previous studies [29,39] and can be explained by the main effect model of social support, which states that social support has a generalized beneficial effect [51]. This finding suggests that in-person social support theory can be extended to the online world of social relations. Moreover, given that the data are cross-sectional, it is also possible that adolescents with higher life satisfaction received more social support on social media. The overall OSS was not significantly associated with depression, while the instrumental and informational dimensions of OSS were negatively associated with depression. A recent meta-analysis revealed a small and insignificant correlation between OSS and adolescents’ depression ($r = -0.09$) [22]. As previous studies mainly discussed the overall impact of OSS, the findings of our study highlight the need to further examine the effects of the dimensions of OSS and the individual difference characteristics that moderate such effects. In line with Nick et al. [9], the results of the current study showed that spending more time on social media was associated with a higher level of OSS. The study further revealed that all kinds of OSS were increased with more time spent on social media.

The current study contributes to the current literature in several ways. First, to the best of our knowledge, this is the first psychometrically sound instrument in Chinese to measure OSS received from social media. The study expanded the utility of the OSSS to the Chinese context. The scale can serve as an important tool for assessing differences and similarities in OSS across social, cultural, and media contexts. Second, our study demonstrated the replicability of the OSSS in the adolescent population (aged 12–16) by demonstrating its reliability and validity. Third, the study improved the brevity and ease of administration
of the original scale by reducing the form from 40 to 20 items. Overall, the present study highlights the value of the OSSS in measuring OSS. The tool is now readily available for future research on OSS among the Chinese adolescent population and its corresponding antecedents and consequences.

Although the study provides a useful scale for researchers, some limitations must be considered. First, convenience sampling at the school level limits the generalizability to other schools. Thus, future research must use probability sampling strategies to improve representativeness. Second, the present study has a narrowed focus on junior high school students, as early adolescence is a critical period for educating young people on how to use the Internet and media. The OSS-CS’s utility in the Chinese university student sample and the general sample has not yet been examined. Thus, future studies may benefit from extending the investigation to other populations in China and replicating the study in diverse cultural contexts. Third, the study only assessed the association of the OSSS-CS with depression and life satisfaction. Future research can examine the relationship of the OSSS-CS to a wider variety of outcomes, which would help verify the criterion validity of the measure and fully understand the role OSS played in adolescents’ well-being. Finally, the present study indicated the general benefits of OSS; however, the group differences and mechanisms underlying its benefits remain unknown. Furthermore, longitudinal studies are required to identify predictors of the OSSS-CS.

5. Conclusions

This study adapted and validated a short measure of the OSSS and confirmed that it includes four factors, i.e., esteem/emotional support, social companionship, informational support, and instrumental support. Our results suggest that the OSSS-CS has high internal consistency, construct validity, and criterion validity and demonstrated the invariance of the variance-covariance matrices, factor structure, factor loadings, and factor variance across two independent samples. The OSSS-CS can serve as a useful tool in assessing the online occurrence of social support among the Chinese adolescent population.

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Data Availability Statement: The datasets generated during and/or analyzed during the current study are not publicly available due to datasets containing information that could compromise the privacy of research participants. The data that support the findings of this study are available from the corresponding author upon reasonable request.

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Appendix A

Table A1. Online Social Support Scale—Chinese Short Version.

| 1 = Never 2 = Rarely 3 = Sometimes 4 = Often 5 = A lot |
|-------------------------------------------------------|
| 1. In the social media, someone concerned me [People show that they care about me online.] | 1 | 2 | 3 | 4 | 5 |
| 2. In the social media, someone said or did things that make me feel good about myself. | 1 | 2 | 3 | 4 | 5 |
| 3. In the social media, someone encouraged me [People encourage me when I’m online.] | 1 | 2 | 3 | 4 | 5 |
| 4. In the social media, someone paid attention to me [People pay attention to me online.] | 1 | 2 | 3 | 4 | 5 |
| 5. In the social media, someone told me they like the things I say or do. | 1 | 2 | 3 | 4 | 5 |
| 6. In the social media, and I others to do things with other people. | 1 | 2 | 3 | 4 | 5 |
| 7. Someone in the social media spent time with me [People spend time with me online.] | 1 | 2 | 3 | 4 | 5 |
| 8. Someone in the social media and I do fun things with me online. | 1 | 2 | 3 | 4 | 5 |
| 9. In the social media, I belong to groups of people with similar interests. | 1 | 2 | 3 | 4 | 5 |
| 10. In the social media, I feel like I belong. | 1 | 2 | 3 | 4 | 5 |
| 11. In the social media, I people need with helpful information. | 1 | 2 | 3 | 4 | 5 |
| 12. In the social media, I people would tell me where to find help if I needed it. | 1 | 2 | 3 | 4 | 5 |
| 13. In the social media, I people help me learn new things when I’m online. | 1 | 2 | 3 | 4 | 5 |
| 14. In the social media, I make suggestions to me online. | 1 | 2 | 3 | 4 | 5 |
| 15. In the social media, I people tell me things I want to know online. | 1 | 2 | 3 | 4 | 5 |
| 16. In the social media, I people help me with school or work. | 1 | 2 | 3 | 4 | 5 |
| 17. In the social media, I online to find people, I need. | 1 | 2 | 3 | 4 | 5 |
| 18. In the social media, I people help me with causes or events that I think are important. | 1 | 2 | 3 | 4 | 5 |
| 19. In the social media, I people offer me things I need. | 1 | 2 | 3 | 4 | 5 |
| 20. In the social media, I people help me with school or work, I get help from others online. | 1 | 2 | 3 | 4 | 5 |

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