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Fatness-Associated FTO Gene Variant Increases Mortality Independent of Fatness – in Cohorts of Danish Men

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Abstract

Background: The A-allele of the single nucleotide polymorphism (SNP), rs9939609, in the FTO gene is associated with increased fatness. We hypothesized that the SNP is associated with morbidity and mortality through the effect on fatness.

Methodology/Principal Findings: In a population of 362,200 Danish young men, examined for military service between 1943 and 1977, all obese (BMI≥31.0 kg/m²) and a random 1% sample of the others were identified. In 1992–94, at an average age of 46 years, 752 of the obese and 876 of the others were re-examined, including measurements of weight, fat mass, height, and waist circumference, and DNA sampling. Hospitalization and death occurring during the following median 13.5 years were ascertained by linkage to national registers. Cox regression analyses were performed using a dominant effect model (TT vs. TA or AA). In total 205 men died. Mortality was 42% lower (p = 0.001) with the TT genotype than in A-allele carriers. This phenomenon was observed in both the obese and the randomly sampled cohort when analysed separately. Adjustment for fatness covariates attenuated the association only slightly. Exploratory analyses of cause-specific mortality and morbidity prior to death suggested a general protective effect of the TT genotype, whereas there were only weak associations with disease incidence, except for diseases of the nervous system.

Conclusion: Independent of fatness, the A-allele of the FTO SNP appears to increase mortality of a magnitude similar to smoking, but without a particular underlying disease pattern barring an increase in the risk of diseases of the nervous system.

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Competing Interests: TIAS has various collaborations with industries on obesity research as indicated on the website (http://www.ipm.hosp.dk/person/tias.htm)

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Introduction

Various studies have shown that a series of common single nucleotide polymorphisms (SNPs) in the first intron of the FTO gene is associated with higher body-mass index (BMI) and risk of obesity in European cohorts [1–4]. These SNPs are in tight linkage disequilibrium so studies of only one of them will convey the effect. Since fatness, and especially abdominal fatness, is associated with increased mortality [5–7], we hypothesized that the FTO SNP, rs9939609, T/A with a minor A-allele frequency in Caucasians of 0.45, is associated with mortality through its association with fatness-related diseases (GenBank accession no.: NT_010498). We therefore investigated the association between this SNP and morbidity and mortality with and without adjustment for fatness.

Methods

Study population

Among 362,200 Caucasian men examined at the mean age of 20 years at the draft boards in Copenhagen and its surroundings during 1943–77, a randomly selected group of one in every hundred men (n = 3,601) and all obese men (n = 1,930) were manually identified (referred to as survey S-20) [8;9]. Obesity was defined as 35% overweight relative to a local standard in use at the time, corresponding to a BMI≥31.0 kg/m², which proved to be above the 99th percentile. All obese and half of the random sample, still living in the region, were invited to a follow-up survey in 1992–94. Due to the long sampling period (1943–1977) the age range at follow-up was wide, resulting in a mean age of 46 years (referred to as survey S-46) [4;10]. The participation rate was 50.5% among the obese and 64.4% among the randomly sampled, equivalent to 795 obese and 920 randomly sampled men attending the examination (figure 1), which included sampling of blood, from which DNA was extracted. Genotyping of FTO rs9939609 (Taqman allelic discrimination; KBiosciences, Cambridge, UK) was successful in 96% of the samples - 752 obese and 876 randomly sampled men - with an error rate of 0.72% calculated from 553 duplicate samples in men that additionally had blood drawn at a mean age of 49 years (referred to as S-49). In case of genotype discrepancy, the genotype from S-49 was used, since the
quality of these samples was judged to be of highest quality. Population stratification may occur due to differences in allele frequencies between the obese and randomly sampled men owing to systematic differences in ancestry rather than association of genes with the response variable. However, our cohort study design of Danish Caucasian men where the randomly sampled men come from the same population in which the obese men were identified effectively prevents population stratification. This is reflected in the genotype distributions of the sample, which obeyed Hardy-Weinberg equilibrium (p = 0.27). The attrition of the groups over time implied that the segment of subjects that were followed up and genotyped represent a population of approximately 175,000 men (represented by the 0.5% random sample of 876 men (200*876 = 175,200) and the originally obese men in this population. The median BMI at time of draft board examination was 21.3 (range: 15.7–30.9) among randomly sampled men included in the present study compared to a median BMI of 21.4 (range: 15.7–30.9) for randomly sampled men excluded from this study (for any reason). Similarly, for obese men included in the present study the median BMI at draft board examination was 32.5 (range: 31.0–31.8) versus 32.8 (range: 31.0–45.9) among obese men not included in the present study (for any reason).

Objective measures of weight and height were obtained at S-20 and S-46, and additionally, waist circumference was measured at S-46. BMI was calculated as weight in kilograms divided by the square of height in meters. Total body fat mass was assessed by bioimpedance at S-46, and fat body-mass index (fat-BMI) was calculated as weight of body fat mass in kilograms divided by the square of height in meters. Bioimpedance has proved to be reasonable accurate for assessing body composition (fat mass and fat free mass) in cohort studies [11].

The Danish Civil Registration System was established on April 2, 1968 [12]. All Danish residents alive or born thereafter have been assigned a unique identification number, referred to as their CPR number. Information on vital status and occurrence of disease was obtained by linking the subjects’ CPR number to the Civil Registration System and the National Hospital Discharge Register [13], respectively, until October 31, 2007. Information on causes of death was obtained by linkage to the National Cause of Death Register [14] until 2007. Using an explorative approach, the sequences of diseases that had contributed to death according to the death certificate were investigated. In case of co-morbidity, a subject was only allowed to appear in each disease category once, but in several disease categories simultaneously. An additional approach was undertaken to investigate the occurrence of disease in the year preceding death in men who died during follow-up. Diseases occurring in this year convey supplementary information on diseases that may have contributed to death. The incidence of

* Men who could not be identified or had moved out of the region were not invited to participate in the survey S-46

Figure 1. Participation flow chart from draft board examination.
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Statistical analyses

Our study sample represents the entire BMI distribution through the random sample and with the right end of the distribution being richly represented. Therefore, the data actually consists of a cohort with two nested subcohorts based on the different sampling fraction across the range of BMI at draft board examination. This design implies that the data can be analysed as combined cohort data [13].

For each Cox proportional-hazards model the subjects were followed from date of entry into S-46 until date of death, censoring due to loss to follow-up, emigration, or end of follow-up at October 31, 2007, in the Civil Registration System. The largest group (the carriers of the A-allele) was used as reference, in order to obtain the most precise estimates. The underlying time-scale was age, with delayed entry from age at S-46, which maximizes comparability by age [16].

Results

Out of the total of 1,628 men included in the analyses, 205 men died during the median 13.5 years of follow-up. The genotype distribution differed as expected between the originally obese men (TT, 25.5%; TA, 47.6%; AA, 26.9%) and the randomly sampled men (TT, 32.7%; TA, 50.6%, AA, 16.8%), corresponding to a minor allele frequency of 0.51 in the obese group, 0.42 in the random sampled group and 0.46 combined. The obese were slightly younger than the other men due to the increasing prevalence of obesity during the recruitment period (Table 1). The odds ratio for having a BMI≥31.0 at time of draft board examination was 2.05 (95% CI, 1.55–2.71; p<5*10^{-7}) for the AA genotype versus the TT genotype.

All-cause mortality

The mortality in the obese group was, as expected, greater than in the randomly selected group, with a HR of 2.03 (95% CI, 1.50–2.71).

Table 1. Distribution of age and anthropometrics given as median and range of 1,628 Danish men according to genotype for FTO rs9939609

|        | TT N = 478 | TA N = 801 | AA N = 349 |
|--------|------------|------------|------------|
| No of subjects | Obese 192 | 358 | 202 |
|           | Random sample 286 | 443 | 147 |
| No of deaths | Obese 14 | 54 | 30 |
|           | Random sample 26 | 63 | 18 |
| Age (yr) S-46 | Obese 41 (34–66) | 43 (33–73) | 41 (33–75) |
|           | Random sample 46 (34–71) | 47 (34–73) | 46 (34–67) |
| BMI (kg/m²) S-20 | Obese 32.5 (31.0–42.8) | 32.5 (31.0–45.4) | 32.7 (31.0–51.8) |
|           | Random sample 21.3 (15.9–29.4) | 21.2 (15.7–30.9) | 21.4 (16.3–29.7) |
| BMI (kg/m²) S-46 | Obese 34.6 (24.3–63.7) | 34.9 (19.9–51.5) | 35.6 (25.2–59.6) |
|           | Random sample 25.7 (18.0–36.5) | 25.8 (16.2–45.1) | 25.9 (19.4–41.9) |
| Waist (cm) S-46 | Obese 115.5 (79.5–183.0) | 116.8 (84.0–160.0) | 117.0 (81.0–171.5) |
|           | Random sample 92.2 (70.0–130.0) | 93.0 (66.0–131.0) | 93.2 (58.0–139.0) |
| Fat-BMI (kg/m²) S-46 | Obese 11.8 (4.9–32.1) | 12.0 (3.4–24.3) | 12.4 (4.7–26.7) |
|           | Random sample 6.3 (1.2–13.3) | 6.4 (0.3–19.2) | 6.2 (1.5–16.6) |

Abbreviations: BMI, body mass index; fat-BMI, fat mass (kg)/height² (m²); N, number of subjects; S-20, draftee population of median age 20; S-46, follow-up survey at median age 46

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A dominant effect of the A-allele (TT genotype versus TA and AA genotypes) gave the best fit to the mortality data as illustrated in Table 2. This was confirmed by likelihood ratio tests (LRT) where an additive effect of the gene (LRT, 2.65; P = 0.10), a dominant effect of the gene (LRT, 0.45; P = 0.50) and a recessive effect of the gene (LRT, 10.77; P = 0.001) were tested. A decreased mortality in men with the TT genotype compared to A-allele carriers was observed, both among obese and the randomly sampled, and since there was no interaction, a pooled estimate was made; HR, 0.58; 95% CI, 0.41–0.82 (Table 3). The association was attenuated only slightly and remained significant after adjustment for BMI at S-46, waist circumference and BMI at S-46, weight history (i.e. BMI at S-20 and S-46), and fat-BMI at S-46 (Table 3).

Exploratory study of disease pattern

Fat-BMI apparently had the strongest effect on the association between the genotypes and all-cause mortality (Table 3), but since data on fat-BMI was missing for 103 subjects, the cause-specific analyses were instead adjusted for BMI at S-20 and S-46. The results were essentially the same when fat-BMI was included as a covariate instead of BMI at S-20 and S-46. Table 4 shows the disease groups explored (ICD codes are available on request).

Circulatory diseases and cancers accounted for the highest numbers of prevalent diseases at death, irrespective of genotype. For the majority of the disease categories, the estimates were lower for the TT genotype than for the A-allele carriers, although the estimates were not significantly different from 1.00 (Table 4). Remarkably, no individuals with the TT genotype died from or with diseases of the nervous system, whereas ten such cases were observed in the A-allele carriers.

Regarding diseases occurring during the year preceding death, the HR's were lower for the TT genotype (vs. TA and AA genotypes) in all categories, though not significantly different from 1.00 (Table 4).

The HRs for incidence of diseases during the follow-up after genotyping were in general very close to 1.00, indicating no difference in risk between the genotypes. However, a significantly lower risk was observed for endocrine diseases in TT versus A-allele carriers (HR, 0.73; 95% CI, 0.55–0.96) as well as for diseases of the nervous system (HR, 0.57; 95% CI, 0.40–0.82).

Discussion

The present findings show a strong fatness-independent association between the FTO SNP, rs9939609, and all-cause mortality in middle-aged men of a magnitude similar to that of smoking [18], with a significantly decreased risk of dying in the TT genotype group compared with A-allele carriers. This was observed both in the random sample of the 362,200 Danish draftees and in the selected sub-sample of the most obese men in this population. Our a priori hypothesis – that the A-allele of the FTO SNP was associated with increased mortality and morbidity primarily through an effect on fatness – was thus not confirmed, as our results showed only a slight attenuation of the HR by adjusting for the fatness-covariates.

We observed that men with the AA genotype had a 2.05 increased odds of obesity compared with the TT genotype, which is higher than reported in previous studies [1;19;20]. This difference in odds ratios is most likely due to differences between the examined cohorts. Our cohort is characterized by a massive enrichment of the right tail of the BMI distribution, making it possible to demonstrate a stronger association compared to already published studies examining the effects of FTO on fatness. Men with the TT genotype had decreased mortality from most causes assessed. Our results suggest that the FTO SNP may affect particularly the nervous system, however, the number of cases with disease of the nervous system was small and could not alone account for the excess mortality observed among A-allele carriers.

In addition to death certificates providing the presumed sequence of diagnoses leading to death as well as other prevalent diseases, we also examined inpatient registry data on diseases occurring close to time of death, thereby obtaining the best

### Table 2. FTO rs9939609 genotype distribution and all-cause mortality in 1,628 Danish men

|                | TT    | TA    | AA    |
|----------------|-------|-------|-------|
|                | HR    | 95% CI| HR    | 95% CI| HR    | 95% CI|
| Obese          | 0.52  | 0.29–0.94 | 1.00  | 0.86–1.66 | 0.06  |        |
| Random sample  | 0.63  | 0.40–0.99 | 1.00  | 0.60–1.73 | 0.02  |        |
| Combined       | 0.59  | 0.41–0.84 | 1.00  | 0.75–1.48 | 0.00  |        |

Cox proportional hazards analysis with age as the underlying time scale and delayed entrance at age of BMI at S-20 and S-46. Abbreviations: HR, Hazard ratio; CI, Confidence interval.

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### Table 3. FTO rs9939609 (genotype TT vs. TA and AA) and all-cause mortality in 1,628 Danish men

|                | Obese | Random sample | Pooled |
|----------------|-------|---------------|--------|
|                | N     | HR 95% CIs    | P value| N     | HR 95% CIs    | P value| N     | HR 95% CIs    | P value|
| Crude          | 1,628 | 0.51 0.29–0.90 | 0.02  | 0.62 0.40–0.97 | 0.04  | 0.58* 0.41–0.82 | 0.001 |
| Adjusted for BMI S-46 | 1,626 | 0.52 0.29–0.91 | 0.02  | 0.61 0.39–0.95 | 0.03  | 0.58* 0.41–0.82 | 0.002 |
| Adjusted for BMI and waist S-46 | 1,624 | 0.51 0.29–0.91 | 0.02  | 0.59 0.38–0.92 | 0.02  | 0.57* 0.40–0.81 | 0.002 |
| Adjusted for BMI S-20 and S-46 | 1,626 | 0.53 0.30–0.94 | 0.03  | 0.61 0.39–0.96 | 0.03  | 0.59* 0.41–0.83 | 0.003 |
| Adjusted for fat-BMI S-46 | 1,523 | 0.59 0.33–1.05 | 0.07  | 0.64 0.41–1.01 | 0.06  | 0.63* 0.44–0.90 | 0.01  |

Cox proportional hazards analysis with age as the underlying time scale and delayed entrance at age of BMI at S-20 and S-46. There was no interaction between the genotypes and the original sampling variable (Random sample vs. obese).

Abbreviations: BMI, body mass index; fat-BMI, fat mass (kg)/height2 (m2); N, number of subjects with complete covariate information; HR, Hazard ratio; CI, Confidence interval; S-20, draftee population of median age 20; S-46, follow-up survey at median age 46.

p < 0.05 for the likelihood ratio test. (Likelihood ratio tests were used to estimate whether a model including both fatness-covariates and FTO rs9939609 fitted the data better than a model including only fatness-covariates).

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Table 4. FTO rs9939609 (genotype TT vs. TA and AA) and cause-specific mortality and morbidity in 1,628 Danish men adjusted for BMI S-20 and S-46

| Category                        | TT  | HR     | 95% Cls | No | No  | AA    | HR     | 95% Cls |
|--------------------------------|-----|--------|---------|----|-----|-------|--------|---------|
| **Infectious diseases**        |     |        |         |    |     |       |        |         |
| Prevalent at time of death     | 3   | 1.17   | 0.30–4.60 | 7  | 1.00 |
| Incident one year prior to death | 4  | 0.55   | 0.18–1.63 | 17 | 1.00 |
| First incidence                | 37  | 0.95   | 0.65–1.39 | 107| 1.00 |
| **Cancer**                     |     |        |         |    |     |       |        |         |
| Prevalent at time of death     | 13  | 0.84   | 0.43–1.66 | 51 | 1.00 |
| Incident one year prior to death | 14 | 0.68   | 0.38–1.25 | 48 | 1.00 |
| First incidence                | 41  | 0.84   | 0.58–1.20 | 114| 1.00 |
| **Haematological diseases**    |     |        |         |    |     |       |        |         |
| Prevalent at time of death     | 0   | 0.00   | -       | 3  | 1.00 |
| Incident one year prior to death | 3  | 0.55   | 0.16–1.92 | 14 | 1.00 |
| First incidence                | 11  | 0.72   | 0.37–1.41 | 43 | 1.00 |
| **Endocrine diseases**         |     |        |         |    |     |       |        |         |
| Prevalent at time of death     | 7   | 0.62   | 0.28–1.41 | 35 | 1.00 |
| Incident one year prior to death | 11 | 0.73   | 0.37–1.43 | 40 | 1.00 |
| First incidence                | 64  | 0.73   | 0.55–0.96 | 238| 1.00*|
| **Mental diseases**            |     |        |         |    |     |       |        |         |
| Prevalent at time of death     | 1   | 0.20   | 0.03–1.52 | 12 | 1.00*|
| Incident one year prior to death | 2  | 0.33   | 0.07–1.48 | 13 | 1.00 |
| First incidence                | 23  | 0.92   | 0.56–1.49 | 63 | 1.00 |
| **Diseases of the nervous system** |    |        |         |    |     |       |        |         |
| Prevalent at time of death     | 0   | 0.00   | -       | 10 | 1.00 |
| Incident one year prior to death | 2  | 0.27   | 0.06–1.17 | 19 | 1.00*|
| First incidence                | 38  | 0.57   | 0.40–0.81 | 183| 1.00*|
| **Circulatory diseases**       |     |        |         |    |     |       |        |         |
| Prevalent at time of death     | 20  | 0.70   | 0.42–1.15 | 74 | 1.00 |
| Incident one year prior to death | 20 | 0.73   | 0.44–1.21 | 70 | 1.00 |
| First incidence                | 122 | 0.90   | 0.73–1.11 | 360| 1.00 |
| **Respiratory diseases**       |     |        |         |    |     |       |        |         |
| Prevalent at time of death     | 6   | 0.54   | 0.22–1.33 | 26 | 1.00 |
| Incident one year prior to death | 9  | 0.49   | 0.24–1.02 | 44 | 1.00*|
| First incidence                | 50  | 0.83   | 0.60–1.15 | 157| 1.00 |
| **Diseases of the digestive system** |  |        |         |    |     |       |        |         |
| Prevalent at time of death     | 7   | 0.72   | 0.31–1.68 | 23 | 1.00 |
| Incident one year prior to death | 9  | 0.85   | 0.39–1.81 | 27 | 1.00 |
| First incidence                | 92  | 0.94   | 0.74–1.20 | 246| 1.00 |
| **Diseases of the locomotion system** | |        |         |    |     |       |        |         |
| Prevalent at time of death     | 0   | 0.00   | -       | 5  | 1.00 |
| Incident one year prior to death | 7  | 0.94   | 0.39–2.26 | 20 | 1.00 |
| First incidence                | 152 | 1.12   | 0.92–1.36 | 347| 1.00 |
| **Diseases of the kidney/urinary system** | |        |         |    |     |       |        |         |
| Prevalent at time of death     | 2   | 0.59   | 0.12–2.82 | 8  | 1.00 |
| Incident one year prior to death | 6  | 0.66   | 0.26–1.63 | 22 | 1.00 |
| First incidence                | 59  | 0.86   | 0.64–1.15 | 181| 1.00 |
| **Accidents/external causes**  |     |        |         |    |     |       |        |         |
| Prevalent at time of death     | 3   | 1.07   | 0.27–4.31 | 6  | 1.00 |
| Incident one year prior to death | 13 | 0.80   | 0.43–1.51 | 40 | 1.00 |
| First incidence                | 213 | 1.06   | 0.90–1.25 | 512| 1.00 |

**Table 4. Cont.**

| Category                        | TT  | HR     | 95% Cls | No | No  | AA    | HR     | 95% Cls |
|--------------------------------|-----|--------|---------|----|-----|-------|--------|---------|
| Other/unspecified diseases      |     |        |         |    |     |       |        |         |
| Prevalent at time of death     | 5   | 0.68   | 0.25–1.86 | 17 | 1.00 |
| Incident one year prior to death | 19  | 0.56   | 0.34–0.92 | 83 | 1.00*|
| First incidence                | 143 | 0.96   | 0.79–1.16 | 380| 1.00 |

Cox proportional hazards analysis with age as the underlying time scale and delayed entrance at age at time of blood sampling (S-46). Cause of death information was not available in 21 men who died after December 31, 2006. *p < 0.05 for the likelihood ratio test (Likelihood ratio tests were used to estimate whether a model including both BMI S-20, BMI S-46 and FTO rs9939609 fitted the data better than a model including only BMI S-20 and BMI S-46) *p < 0.05 for interaction between genotype and the original sampling variable (random vs. obese); the interaction term is included in the model.

The category other/unspecified diseases includes benign neoplasms, diseases of the eye, skin and subcutaneous tissue, congenital malformation of the urinary system and symptoms, signs and abnormal clinical and laboratory findings, not otherwise classified.

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There are several strengths of this study, but also limitations. It is based on a unique, well-defined large background study population of white men with the obese and random sample derived from the same quite homogenous population, hence eliminating population stratification. The complete sampling of an obese group facilitates valid estimation of the associations between FTO SNP, fatness and morbidity and mortality, and hence also coverage of diagnoses contributing to death. As for prevalent diseases at death, assessment of these diseases showed lower risk estimates for almost all disease categories among individuals with the TT genotype. When analyzing incidence of diseases emerging earlier on (but still after genotyping), the TT genotype was less protective, hence suggesting that the genotype affects the subject’s general ability to cope with disease once present rather than affecting susceptibility to developing disease. However, for diseases of the nervous and endocrine systems, a decreased occurrence was observed already at this stage.

The FTO gene is of ancient origin and conserved in vertebrates and marine algae [21;22]. Bioinformatics analyses suggests that FTO encodes a DNA demethylase, up-regulated by feeding and down-regulated by fasting [23] and studies suggest that the SNP influences cerebro-cortical insulin sensitivity [24] and lipolytic activity in adipose tissue of healthy individuals [25]. Our findings suggest that the FTO SNP may have much wider implications than the effect on fatness and the effects derived hereof. The observation of an increased mortality among A-allele carriers over the broad range of disease categories may translate into the hypothesis that the cellular effects of the SNP affects most organs, and particularly their ability to maintain or regenerate proper function when afflicted by various diseases. The positive association between A-allele carriers and disease of the nervous system also needs further investigation. It may be related to the high expression of FTO in fetal and adult brain tissues [1;2;21;23]. Due to the limited numbers of cases and the multiple testing, we emphasize that our findings are purely hypothesis generating. To test these hypothesis further research is warranted, both at the level of mapping disease patterns at a more detailed level than the organ specific disease categories performed here, and through further examination of the functional impact of the FTO SNP. Since the SNP is intronic it may exert functional effects through altered expression of FTO, or the observed effects may putatively be mediated via another gene.

There are several strengths of this study, but also limitations. It is based on a unique, well-defined large background study population of white men with the obese and random sample derived from the same quite homogenous population, hence eliminating population stratification. The complete sampling of an obese group facilitates valid estimation of the associations between the FTO SNP, fatness and morbidity and mortality, and hence also of fatness-independent associations. The inclusion of a randomly selected group gave us the possibility to reproduce our estimates as
well as to study the broad range of BMI when combining the groups. The accessibility of complete register-based data on morbidity and mortality is an asset of the study, although the clinical routine procedure generating the information requires cautious interpretation. The sample size may seem small for a genetic association study, but this apparent limitation is counteracted by the fact that the random sample is representative of 175,000 men, among whom the obese participants represent the most extreme range of fatness phenotypes in the same population, except for the possible selective attrition before 8-46, where the genotyping was performed. The obese group exhibited the expected excess mortality relative to the randomly sampled group. The finding that circulatory diseases and cancers were the categories with the largest number of cases, further confirmed that the sample was representative of the general male population, since these are listed as main causes of death in the Danish male population [26]. Further, we investigated morbidity and mortality in middle-aged men, hence, we cannot conclude on the FTO SNP’s impact on disease-risk in childhood or young adulthood. The major limitation of the study is that the number of deaths is too small, and the information about the diseases leading to death too crude, to conduct a thorough disease-specific genotype-phenotype investigation. Given these limitations, analyses of the reasons for the excess mortality obviously need to be explored in other populations of both men and women.

In conclusion, the FTO SNP, rs9939609, may have a fatness-independent impact on mortality and morbidity leading to death, with a general lower risk in the TT genotype compared with A-allele carriers.

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Author Contributions
Analyzed the data: EZ CH. Wrote the paper: EZ TIAS TJ. Reviewed and validated data and final analyses: TJ SIJK TB CH THP TH OP TIAS. Contributed to the interpretation of data and revision of the report, and approved the final version: TJ SIJK TB CH TPH TH OP TIAS. Initiated the original and the current study: TIAS.

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