New insights on the $z$–$\alpha$ correlation from complete radio samples

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ABSTRACT

The existence of a correlation between observed radio spectral index and redshift has long been used as a method for selecting high-redshift radio galaxy candidates. We use nine highly spectroscopically complete radio samples, selected at different frequencies and flux limits, to determine the efficiency of this method and compare consistently observed correlations between spectral index ($\alpha$), luminosity ($P$), linear size ($D$) and redshift ($z$) in our samples. We observe a weak correlation between $z$ and $\alpha$ which remains even when Malmquist bias is removed. The strength of the $z$–$\alpha$ correlation is dependent on both the $k$-correction and sample selection frequency, in addition to the frequency at which $\alpha$ is measured, and consistent results for both high- and low-frequency-selected samples are only seen if analysis is restricted to just extended radio galaxies. This fits with the popular interpretation that the spectra steepen with $z$ because the radio lobes work against a denser intergalactic medium environment as $z$ increases, out to $z \sim 2$–3. However, we also note that the majority of sources known at $z > 4$ are very compact and often display a negatively curved or peaked spectrum, indicative of youth or merger activity, and therefore the low-frequency radio spectrum as a whole should be determined; this is something for which the new LOw Frequency ARray will be crucial. We quantify both the efficiency and the completeness of various techniques used to select high-$z$ radio candidates. A steep-spectrum cut applied to low-frequency-selected samples can more than double the fraction of high-$z$ sources, but at a cost of excluding over half of the high-$z$ sources present in the original sample. An angular size cut is an almost as equally effective radio-based method as a steep-spectrum cut for maximizing the high-$z$ content of large radio samples, and works for both high- and low-frequency-selected samples. In multiwavelength data, selection first of infrared-faint radio sources remains by far the most efficient method of selecting high-$z$ sources. We present a simple method for selecting high-$z$ radio sources, based purely on combining their observed radio properties of $\alpha$ and angular size, with the addition of the $K$-band magnitude if available.

Key words: galaxies: active – galaxies: evolution – galaxies: high-redshift – radio continuum: galaxies.

1 INTRODUCTION

Vast amounts of multifrequency radio data at long wavelengths will soon begin to flow from next-generation radio instruments such as the LOw Frequency ARray (LOFAR) and eventually the Square Kilometre Array (SKA). With this, opportunities will arise for studying some of the earliest radio sources in the Universe, their environments and their evolution over cosmic time. There is also the tantalizing possibility of studying conditions within the epoch of reionization itself through high-$z$ radio sources: if sufficiently bright radio sources can be found at redshifts greater than 6.5, it should be possible to measure absorption signatures of neutral hydrogen, and hence trace changes in the ionization state of the Universe with cosmic time (e.g. Carilli et al. 2007; Meiksin 2011).

The existence of a correlation between redshift and observed spectral index $\alpha$ (where $S \propto \nu^\alpha$) for powerful radio sources was first suggested by Tielens, Miley & Willis (1979) who found that the introduction of a steep spectral index cut led to an increasing fraction of sources without optical counterparts identified on the First Palomar Sky Survey ($R < 20$) plates. Blundell, Rawlings & Willott (1999) used the combination of complete samples from the Third Cambridge Revised Revised (3CRR; Laing, Riley & Longair 1983),

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6CE (Rawlings, Eales & Lacy 2001) and Seventh Cambridge Redshift Survey (7CRS; Lacy et al. 1999; Willott et al. 2003) surveys to confirm that the high-frequency (5 GHz) rest-frame spectral index correlates with redshift, but showed that this correlation is weaker for spectral index measured at lower rest-frame frequencies. Since then, there have been many surveys designed to pick out only ultrasteep spectrum (USS) sources for further infrared imaging and spectroscopic follow-up (e.g. Röttgering et al. 1996; De Breuck et al. 2000), having varying degrees of success selecting high-z sources. Many of these have additional selection criteria such as small angular size and faint infrared magnitude applied after the USS cut, which makes it difficult to determine the extent to which the USS cut is responsible for selecting high-z sources.

Klamer et al. (2006) present a sample of USS-selected galaxies selected from the Sydney University Molonglo Sky Survey (SUMSS), and discuss the apparent mechanisms for the z–α correlation in detail. They dismiss the possibility of k-corrections being the cause of the observed steepening radio spectra, given that the majority of their radio spectra show no evidence for curvature. They suggest that enhanced spectral aging due to inverse-Compton losses against the cosmic microwave background at high redshift is the most likely origin for the observed z–α correlation, or that it may arise due to an intrinsic relation between low-frequency α and radio luminosity coupled to a Malmquist bias. Following on from this work, Bryant et al. (2009) compare the median redshift of several complete samples with the median redshift obtained from USS-selected samples. They find it to be lower in complete samples, and argue that this is strong evidence for the efficiency of the USS technique. However, whilst USS samples clearly do select higher redshift sources, sample comparisons between USS and complete samples are not ideal for either optimizing or quantifying the efficiency of the technique. The most rigorous approach would be to apply selection criterion to complete samples with significant numbers of sources at the highest redshifts currently known, and quantify the number of high-redshift sources included/missed.

Despite the wide use of the z–α technique to select high-redshift galaxies, there has been very little work on quantifying the efficiency. Pedani (2003) states that, for the first time, they present the true quantitative searching efficiency for high-z radio galaxies using a sample selected from the Molonglo Reference Catalogue (MRC). They utilize 225 sources with full redshift information from this sample to measure the efficiency of optical, USS and size selection. They find that the efficiency (defined as the fraction of z > 2 sources in the recovered sample) of the USS criterion alone is 0.33, increasing to a maximum of ~0.59 in combination with an optical cut. However, their 225 source sample is not complete, being composed of only objects with redshift information amongst the complete MRC 1 Jy radio sample of 446 sources. They argue that the redshift-complete subsample is representative of the full sample, as both contain similar proportions of USS sources. However, they also note that the median magnitude of the subset of galaxies without redshifts is fainter than that for those included. This means, first, that there is an optical magnitude bias towards brighter magnitudes in the analysed sample and, secondly, that the work is based on the implicit assumption that the USS criterion is more important than optical magnitude in selecting high-redshift candidates. With 50 per cent of the sample not analysed, and at fainter optical magnitudes, the redshift incompleteness towards higher redshift cannot be quantified, and this could be substantial.

Potential biases such as these are common in the literature, due to the difficulty and expense of building spectroscopically complete radio samples. As such, any attempt to use large collections of radio data available in the literature to investigate evolution of basic radio properties is not valid, despite the large number statistics, without clearly defined and well-understood selection criteria. For example, recent work by Khabibullina & Verkhodanov (2009) uses a large sample of 2442 radio galaxies with measured redshifts selected from large publicly available radio source catalogues. They determine the dependence of α on z, and select a sample of distant objects using this relation. Crucially, however, as they note, the samples they use are not complete in any sense, and some of the largest high-z radio source samples with radio spectra publicly available are ones with a USS criterion applied (e.g. De Breuck et al. 2000) which will irrevocably bias spectral index studies of any sample constructed from them.

In summary, although the existence of the so-called z–α correlation has been known for some time, there has been little attempt to quantify the strength of this consistently across a wide range of spectroscopically complete samples at different selection frequencies, and measure the resultant efficiency of using a USS α cut-off in order to isolate high-z candidates. In this study, we address these shortcomings, thus providing a vital tool for the design of further high-z source searches from upcoming radio surveys by new survey instruments, e.g. LOFAR. This work builds significantly on current knowledge in five ways.

(i) We use nine highly spectroscopically complete and unbiased radio source samples. Most have a spectroscopic completeness in the range 80–100 per cent, and robust redshift estimates (e.g. photometric or based on the K–z relation) are available for the vast majority of the remaining sources, such that all samples are at least 95 per cent redshift complete.

(ii) We use new radio data from the Combined EIS-NVSS Survey of Radio Sources (CENSORS) radio sample (Best et al. 2003), which contains a large number of sources with z > 2, improving high-redshift statistics.

(iii) Selection frequency effects are fully explored – four samples are selected at frequencies below 200 MHz, and five at 1.4 GHz.

(iv) The samples have a wide range of flux density limits, so that correlations such as the P–α and z–α relations may be safely disentangled.

(v) We also consider radio linear size (D) in order to investigate its role in selecting high-redshift sources.

The layout of this paper is as follows. In Section 2, we describe the complete radio samples used in this study. In Section 3, we give a brief summary of radio source properties and sample selection effects. In Section 4, we investigate observable trends and employ principal component analysis (PCA) to identify fundamental correlations in the P–D–α–z parameter space for various collections of samples. In Section 5, we attempt to fit various functional forms for α to the observed data, and identify large intrinsic scatter in α, not dependent on P, z or D. In Section 6, we discuss the physical origins of the observed z–α correlation, and finally in Section 7 we discuss the implications of our findings in the search for the highest redshift radio galaxies, and use complete samples to explore the efficiency of often used techniques in the literature to find these.

Throughout this work, a Λ cold dark matter cosmology is assumed, ΩΛ = 0.7, ΩM = 0.3 and H0 = 70 km s⁻¹ Mpc⁻¹, and magnitudes are in the Vega system.

2 COMPLETE RADIO SAMPLES SELECTION

We want to quantify the z–α correlation at a wide range of frequencies and flux density limits, determine to what extent this is an
Table 1. Details of the complete samples used in this study. Note that the spectral indices for CoNFIG are taken from Gendre et al. (2010), from a linear fit to flux densities between 1.4 GHz and 178 MHz, rather than a two-point spectral index.

| Survey | Selection ν (MHz) | No. of sources | Sky area (sr) | Flux limit (Jy) | Per cent z_{spect} | Per cent no z | α range (MHz) |
|--------|------------------|----------------|--------------|----------------|-------------------|---------------|---------------|
| 3CRR   | 178              | 173            | 4.239        | $S_{178} > 10.9$ | 100               | 0.0           | 750–151       |
| 6CE    | 151              | 58             | 0.102        | $2.00 < S_{151} < 3.93$ | 97              | 3.0           | 1400–151      |
| 7Cl    | 151              | 37             | 0.0061       | $S_{151} > 0.51$   | 90               | 0.0           | 1400–151      |
| 7CII   | 151              | 37             | 0.0069       | $S_{151} > 0.48$   | 90               | 0.0           | 1400–151      |
| TOOT00 | 151              | 47             | 0.0015       | $S_{151} > 0.50$   | 95               | 0.0           | 1400–151      |
| WP85r  | 1400             | 138            | 9.81         | $S_{1400} > 4$     | 95               | 0.0           | 5000–1400     |
| CoNFIG1| 1400             | 273            | 1.5          | $S_{1400} > 1.3$   | 83               | 3.6           | 1400–178      |
| CoNFIG2r| 1400           | 61             | 0.89         | $1.0 < S_{1400} < 1.3$ | 52              | 4.9           | 1400–178      |
| PSRr   | 1400             | 59             | 0.075        | $S_{1400} > 0.36$  | 61               | 0.0           | 2700–1400     |
| CENSORS| 1400             | 135            | 0.0018       | $S_{1400} > 0.0072$| 78              | 3.7           | 1400–325      |
| Hercules| 1400            | 64             | 0.00038      | $S_{1400} > 0.002$ | 66               | 3.0           | 1400–610      |

Figure 1. Radio luminosity versus redshift plane coverage for the low-frequency-selected samples 3CRR, 6CE, 7CRS and TOOT (left) and for the high-frequency-selected samples WP85r, PSRr, CoNFIG 1 and 2r, CENSORS and Hercules (right). Only sources with $α < −0.5$, as used in this study, are shown. The boxes indicate the Malmquist-bias-free sections of the $P$–$z$ plane at high and low frequency, which we use in later analysis; for 151 MHz samples: $\log P = 27.75–29.5$, $z = 0.5–3.5$, and for 1.4 GHz samples: $\log P = 26.25–29.0$, $z = 1.0–4.5$.

intrinsic property of sources (rather than being driven by, for example, a $P$–$α$ correlation) and understand any selection effects present. In order to do this, we collate data from several complete samples already available in the literature: the 3CRR, 6CE, 7CRS and TOOT00 selected at low frequency, and the WP85r, Combined NVSS-FIRST Galaxy catalogue (CoNFIG) regions 1 and 2, PSRr, CENSORS and Hercules samples selected at high frequencies. These samples are described below, summarized in Table 1, and displayed on the $P$–$z$ plane in Fig. 1.

2.1 The 3CRR sample

The 3CRR sample of extragalactic radio sources (Laing et al. 1983) is a complete sample containing all radio sources above 10.9 Jy at 178 MHz in an area of sky covering 4.239 sr. The sample comprises 173 objects in total, and is 100 per cent spectroscopically complete. The data were obtained from the 3CRR catalogue web page maintained online.¹ As this sample is the only low-frequency-selected sample observed at 178 MHz as opposed to 151 MHz, the flux densities are converted to 151 MHz fluxes assuming the spectral indices given in the catalogue. The observed spectral index $α$ is measured between 750 and 178 MHz for this sample.

2.2 The 6CE sample

The 6CE sample is based on an original sample selected by Eales (1985) from the Sixth Cambridge radio survey (6C) comprising 67 radio sources between 2.2 and 4 Jy selected at 151 MHz, over a sky area of 0.102 sr. For this study, we use a reselected, updated version, available online,² the 6CE sample of Rawlings et al. (2001). This consists of all sources with a 151 MHz flux density in the range $2.00 \leq S_{151} \leq 3.93$ Jy in the same 0.102 sr patch of sky. There are 59 sources in total, with all but one having a firm identification, and 56 of the 59 having spectroscopic redshifts. The three sources without spectroscopic redshifts, one is obscured by a bright star and so is excluded from the sample, and the other two have a redshift estimate from the $K$–$z$ relation. Observed spectral indices have been calculated between 1.4 GHz and 151 MHz using 1.4 GHz fluxes obtained from the NRAO VLA Sky Survey (NVSS).

2.3 The 7CRS sample

The 7CRS is composed of three subsamples, 7Cl, 7CII and 7CIII. 7Cl and 7CII are each composed of 37 sources with flux density limits of $S_{151} \geq 0.51$ Jy and $S_{151} \geq 0.48$ Jy, respectively, in the 7C survey and are defined in Willott et al. (2003). The redshifts and

¹ http://3crr.extragalactic.info/
² http://www-astro.physics.ox.ac.uk/sr/6ce.html
linear sizes for the 7CI and 7CII samples are available online from the data of Grimes, Rawlings & Willott (2004). The 7CI sample contains 54 objects with a flux limit of $S_{151} > 0.50$ Jy, detailed in Lacy et al. (1999). We utilize the 7CII data from table 8 in Lacy et al. (1999) to get the luminosities, redshifts and linear sizes of the sample.

The spectral indices for this sample are not yet available in a collective form in the literature. We estimate the observed spectral index by cross-matching 151 MHz fluxes for each source from the 7C 151 MHz catalogue of Hales et al. (2007) with the NVSS (Condon et al. 1998), at 1.4 GHz. We checked all extended sources listed in the 7C Hales catalogue as having separate components, and cross-checked maps at 151 MHz with NVSS maps in order to correctly identify components and catalogue the correct integrated flux for each source. We have also matched the source list with the Texas Survey of Radio Sources (TExAS) 365 MHz/Westerbork Northern Sky Survey (WENSS) 327 MHz surveys, the 5C 408 MHz survey, and finally the VLA Low-Frequency Sky Survey (VLSS) at 74 MHz (Cohen et al. 2007) with the addition of the 38 MHz 8C survey for 7CII (Lacy et al. 1999), all of which are of comparable resolution, for later curvature analysis (see Section 6, and Ker et al., in preparation). We note that it is possible that very extended sources may not have correct fluxes in these catalogues.

7CI and 7CII both have 90 per cent spectroscopic redshift completeness, and 7CIII is 95 per cent complete. The remaining sources in all three subsamples have photometric redshifts estimated from the $K$–$z$ relation.

2.4 The TOOT00 sample

The TOOT00 region (Vardoulaki et al. 2010) is the first complete region of the Tex-Ox-One-Thousand (TOOT) redshift survey of radio sources. This survey selects all sources above 100 mJy in the Cambridge 7C 151 MHz survey, and is designed to be approximately five times fainter than the 7CRS, with much greater numbers. Vardoulaki et al. (2010) present complete radio, near-infrared and spectroscopic data or redshift estimates for the first region of the survey, comprising 47 sources. 40 of the radio sources have spectroscopic redshifts, with a further six using a redshift estimated from the $K$–$z$ relation. The final source has a $K$ limit only and we adopt the lower redshift limit as the redshift estimate for this source (the $K$-band data reach sufficient depth to place the source at high-$z$, and hence for the lower redshift limit to be adopted as the redshift estimate with little loss of accuracy). The observed spectral index was calculated for each source using flux data at 151 MHz and 1.4 GHz (NVSS) from Vardoulaki et al. (2010).

2.5 The Wall & Peacock 2.7 GHz sample

The original Wall & Peacock (1985) 2.7 GHz 233 source sample covers 9.81 sr of sky, and includes all radio sources brighter than 2 Jy. The sample now stands at 98 per cent spectroscopically complete (Rigby et al. 2011). In this study, we use the 138 source sample, reselected by Rigby et al. (2011) to be complete at 1.4 GHz to a flux limit of 4 Jy with a spectral index between 5 and 1.4 GHz steeper than $-0.5$. This reselected sample is 97 per cent spectroscopically complete, and the remaining three sources have photometric redshift estimates. We refer to this sample as WP85r.

2.6 The CoNFIG samples

We utilize two complete samples from the CoNFIG regions 1 and 2 (Gendre, Best & Wall 2010).

CoNFIG1 contains 273 sources complete to 1.4 GHz $\geq 1.3$ Jy, and is 83 per cent spectroscopically complete. In CoNFIG 1, 226 sources have spectroscopic redshifts, 37 have photometric redshift estimates and 10 sources (4 per cent) have only lower redshift limits from Sloan Digital Sky Survey (SDSS) $I$-band non-detections. These non-detections are not sufficiently deep to provide a useful constraint on the redshift (the SDSS limiting $I$-band magnitude only constrains each source to $z \geq 1$). However, of these 10 sources, only four have an observed spectral index steeper than $-0.5$, and hence should be included in the analysis (see Section 3). We choose not to include these four sources as the redshift estimate is not reliably constrained, and such a small fraction will have a statistically insignificant effect on the results. All four sources have very different morphological types and spectral indices, so are unlikely to be biased towards any one redshift range.

CoNFIG2 contains 132 sources and is complete between 1.3 and 0.8 Jy at 1.4 GHz (only sources with 1.4 GHz fluxes less than 1.3 Jy were used from CoNFIG2, to ensure no duplication with sources also in CoNFIG1). At fainter flux densities, the redshift completion of CoNFIG2 is relatively low, so we reselect the sample to above 1 Jy at 1.4 GHz, creating a new sample of 61 sources which we refer to as CoNFIG2r. For this revised sample, reselecting to above 1 Jy at 1.4 GHz reduces the proportion of unidentified redshift sources to a negligible number of three. Of these three sources, only two have a spectral index steeper than $-0.5$, and hence should be included in the analysis. As these two have greatly differing spectral indices and morphologies, again they are unlikely to be limited to any one redshift range, and similarly to CoNFIG1 we do not include these two sources in the subsequent analysis.

The observed spectral index is taken from Gendre et al. (2010), calculated using a linear fit to flux data points between 1.4 GHz and 151 MHz. We cross-matched the CoNFIG catalogues with the VLSS and 7C 151 MHz (Hales et al. 2007) catalogue, giving flux data at 74, 151, 365, 408, 1400, 2700 and 5000 MHz for both samples to allow curvature analysis (see Section 6, and Ker et al., in preparation). Frequency coverage for this data set is very good: only 28 sources have no data at 74/151 MHz, and only 15 have no 2.75 GHz data.

2.7 Parkes selected regions

The original Parkes selected regions (Wall, Cole & Milne 1968; Downes et al. 1986; Dunlop et al. 1989) is a complete 178 object sample containing all radio sources brighter than 0.1 Jy over a 0.075 sr sky area at 2.7 GHz. We have reselected the sample at 1.4 GHz $>$ 0.36 Jy as PSR, to which flux limit there are 59 sources with an observed spectral index between 2.7 and 1.4 GHz steeper than $-0.5$. 36 of these sources have spectroscopic redshifts, with the remaining 23 having redshift estimates from the $K$–$z$ relation or photometric spectral fitting.

2.8 The CENSORS sample

The CENSORS sample is a 135 source sample of all radio sources with an NVSS 1.4 GHz flux density greater than 7.2 mJy in a 6 deg$^2$ patch of the sky centred on the ESO Imaging Survey Patch D (Best et al. 2003). The sample currently stands at 96 per cent identified and 78 per cent spectroscopically complete, and is currently one of
Figure 2. A sketch of the contributions of the various components of a typical double radio galaxy to the rest-frame radio regime, and the effects of synchrotron self-absorption, losses and beaming on the observed radio spectrum. Also shown are the observed frequencies at which these features would be observed at redshifts 2 and 6 for comparison.

2.9 The Hercules sample

The Hercules sample is taken from a field in the Leiden-Berkeley Deep Survey, and consists of 64 sources selected to have a flux density greater than 2 mJy at 1.4 GHz (Waddington et al. 2001). The spectroscopic completeness stands at 66 per cent, with 20 sources having photometric redshifts based on the K–z relation, and the final two having a redshift limit estimated from K-band non-detections. At K-band limits of ~19 and above, the non-detections are sufficiently deep that the sources must be at high redshift, and the lower redshift limit can be adopted as the estimated redshift without great loss of accuracy. Observed spectral indices are calculated between 1.4 GHz and 610 MHz.

3 RADIO SOURCE PROPERTIES AND SAMPLE SELECTIONS

Complete radio samples will select very different populations, depending on the frequency at which they are selected and their flux density limit, as different physical contributions dominate at differing rest-frame frequencies. In this study, we utilize samples selected in both the MHz and GHz regimes, representative of existing spectral index studies.

Fig. 2 illustrates the contributing components of a typical extended radio galaxy. If the source is not highly beamed, i.e. not viewed along the jet axis, the emission is dominated at low frequency by synchrotron emission in the radio lobes. Radio lobes typically display a steep-spectrum power-law slope, which can steepen further at higher frequency due to both synchrotron and inverse-Compton losses [see for example Carilli et al. (1991) who analyse the radio spectrum of the well-studied local radio galaxy Cygnus A in depth]. At low frequencies, the lobe spectrum can turn over due to synchrotron self-absorption. The frequency at which this happens depends on both the size and intensity of the emitting component: it occurs at higher frequencies for smaller emitters, leading to the smallest radio sources at sub-kpc size being GHz peaked sources (GPS).

At higher frequencies (above a few GHz), the contribution to the spectrum from the core is often important. Emission from the core is typically flat spectrum, due to the superposition of self-absorbed components of different sizes at the base of the radio jet. If the jet is orientated towards us, it can be Doppler boosted by beaming, and can become dominant at lower frequencies.

As can be seen from Fig. 2, if a sample is selected at a few hundred MHz, up to high redshifts the radio emission will still be probing the lobe-dominated regime, giving a sample of similar, directly comparable sources. However, if a sample is selected at GHz frequencies or above, sources with a significant core component that are orientated such that the jet is aligned along the line of sight towards us (beamed) will be preferentially included, especially at higher redshifts.

It is thought that the observed z–α relation may arise because sources at higher redshift have lobes doing work against a denser medium. Working against a denser medium means there will be less
adiabatic expansion losses, and therefore greater synchrotron losses, with the result that the source is brighter but the radio spectrum steepens faster. However, as shown, in GHz-selected samples, the observed spectral index may alternatively be flattened at the highest redshifts by an increasing contribution of a core component and be less affected by environment.

Although only rest-frame spectral indices should have any direct physical correlation with other observables (observed spectral indices being a good approximation), as far as possible, we utilize a traditional two-point observed spectral index, so as to match the situation most widely encountered in the literature and most simply provided by the observations, e.g., for the selection of USS sources.

There is a strong argument to exclude as far as possible all significantly beamed sources identified in the samples, as not only will their spectral index estimates be distorted (the beamed component generally being flatter spectrum), but also they will be heavily fore-shortened in size. However, as our primary motivation is measuring the efficiency of radio-based correlations in selecting high-redshift radio sources from radio surveys, we must adopt a simple approach to removing these that can be widely applied to blind radio surveys. In most comparable observational studies, a cut of $\alpha = -0.5$ is used as a division, to separate out flat and steep spectrum sources (and indeed such a cut has already been applied in the definition of some of the samples we use). Hence, in order to analyse comparable parts of the radio spectrum, we restrict analysis in this study to only sources with an observed spectral index less than $-0.5$; these will largely be of a similar type (lobe dominated). Sources with a flatter spectral index represent a composite population: as well as quasars sources with an observed spectral index less than 0.5; these will more or less disappear once the $P$–$z$ correlation is removed, whilst the $z$–$\alpha$ and $P$–$\alpha$ decrease in strength.

In Table 2, we present the Spearman rank correlation coefficients for the relations plotted in Fig. 3. Also listed is the two-tailed p-value, which gives an approximate indication of the probability of an uncorrelated system having a Spearman correlation at least as strong as the one calculated from the observed data. The table illustrates several important points. First, it shows that the $P$–$z$ correlation dominates, even when several complete samples are co-added and analysed together; in other words, simply adding several complete samples does not provide sufficient coverage of the $P$–$z$ plane to fully remove the dominant $P$–$z$ correlation. Secondly, it reveals that the $D$–$\alpha$, $D$–$z$ and $z$–$\alpha$ correlations are the strongest observed in both high- and low-frequency-selected samples. The $z$–$\alpha$ correlation is stronger than the $P$–$\alpha$ correlation (correlations between $P$, $\alpha$ and $D$ more or less disappear once the $P$–$z$ correlation is removed). Of particular relevance to this study is the fact that the variation of observed $\alpha$ with size and redshift is relatively weak for both samples.

### 4 Observable Trends

The complete samples detailed previously provide excellent coverage of the $P$–$z$–$\alpha$ parameter space. In Fig. 3, the log $P$, log $D$, $\alpha$ and log $(1 + z)$ planes are plotted, along with best-fitting straight lines to the data. By eye, the data appear to display similar dependencies of spectral index to those reported by Blundell et al. (1999) for the 3CRR, 6CE and 7CRS combined complete radio samples, namely that observed spectral indices steepen with linear size, redshift and radio power (upper panels). Equations of the linear dependencies fitted for spectral index on luminosity, linear size and redshift are given in Tables 5 and 6.

As can be clearly seen from the $P$–$z$ panels in Fig. 3, the use of only one complete sample means that a strong, dominating, $P$–$z$ correlation due to Malmquist bias is present, and this makes disentangling the various dependencies between $P$, $z$, $\alpha$ and $D$ very difficult. The addition of several complete samples mitigates this correlation somewhat, and indeed many previous studies, e.g. Blundell et al. (1999), argue that the combination of several complete samples essentially removes the Malmquist bias.

With the excellent coverage of the $P$–$z$ plane afforded by our nine complete samples, we are able to test if this is indeed the case. We select a Malmquist-bias-free section of the $P$–$z$ plane for both the high- and low-frequency-selected samples, covering a large range in both redshift and radio power (see Fig. 1 for the area definition), and repeat the linear fits (plotted as blue dashed lines in Fig. 3). This utilizes 186 sources in 151 MHz samples and 133 sources in the 1.4 GHz samples (reduced to 56 sources when investigating linear size, as for the 1.4 GHz samples only CENSORS and Hercules have readily available size information). The $D$–$\alpha$ relation appears to increase in strength when the residual $P$–$z$ correlation is removed, whilst the $z$–$\alpha$ and $P$–$\alpha$ decrease in strength.

In Table 2, we present the Spearman rank correlation coefficients for the relations plotted in Fig. 3. Also listed is the two-tailed p-value, which gives an approximate indication of the probability of an uncorrelated system having a Spearman correlation at least as strong as the one calculated from the observed data. The table illustrates several important points. First, it shows that the $P$–$z$ correlation dominates, even when several complete samples are co-added and analysed together; in other words, simply adding several complete samples does not provide sufficient coverage of the $P$–$z$ plane to fully remove the dominant $P$–$z$ correlation. Secondly, it reveals that the $D$–$\alpha$, $D$–$z$ and $z$–$\alpha$ correlations are the strongest observed in both high- and low-frequency-selected samples. The $z$–$\alpha$ correlation is stronger than the $P$–$\alpha$ correlation (correlations between $P$, $\alpha$ and $D$ more or less disappear once the $P$–$z$ correlation is removed). Of particular relevance to this study is the fact that the variation of observed $\alpha$ with size and redshift is relatively weak for both samples.

#### 4.1 Principal Component Analysis

Analysing fully covered sections of the $P$–$z$ plane has shown that relations between $P$, $z$, $\alpha$ and $D$ are strongly coupled to the Malmquist bias. With this in mind, we utilize another statistical test, PCA, which is a technique designed to pick out the intrinsic, dominant linear correlations existing in a multivariate data set, as opposed to secondary correlations arising due to combinations of others (in this case, particularly the Malmquist bias). The method of PCA involves calculating the eigenvalues and eigenvectors, composed of linear combinations of the normalized input parameters, which span the directions of maximum variance in the input data set. These eigenvectors and eigenvalues describe the intrinsic correlations present in the data set (principal components), along with the percentage of the variance in the data that each explain. The results of PCA are most commonly presented in table form, listing each of the principal components, the percentage of the data variance that they explain and the composition of each principal component. Each principal component is composed of a normalized combination of the entered variables, in this case $x$, log $(1 + z)$, log $P$ and log $D$, as $PC = x_1 \alpha + x_2 \log (1 + z) + x_3 \log P + x_4 \log D$, and the final four columns in the table present $x_1$, $x_2$, $x_3$ and $x_4$, showing the relative contributions of each variable for each principal component.

We first look at a low-frequency-selected sample, composed of the 3CRR, 6CE, 7CRS and TOOT00 samples, and perform a PCA analysis (see Table 3, upper). The $P$–$z$ correlation dominates (i.e. the first principal component is along an axis primarily composed of $P$ and $z$), contributing roughly half of the observed variance. A further $\sim 30$ per cent variance is contributed by a $D$–$\alpha$ anticorrelation, whereby sources of larger size have steeper spectra. The
Figure 3. The $\alpha P D$ planes for the 151 MHz (upper six panels) and 1.4 GHz (lower six panels) selected samples. Spectral indices are measured between 1400 and 151 MHz for 6CE, 7CRS, TOOT00 and CoNFIG1 and 2r, 750 and 151 MHz for 3CRR, 5 and 2.7 GHz for WP85r, 2 and 1.4 GHz for PSRr, 1400 and 325 MHz for CENSORS and between 1400 MHz and 610 MHz for Hercules. Only sources with $\alpha$ steeper than $-0.5$ are utilized. The solid green lines indicate the best-fitting straight line to the data. The blue dashed lines indicate the linear fit repeated for a Malmquist-bias-free section of the $P-z$ plane, as defined in Fig. 1. Note that only CENSORS and Hercules are included in the linear size figures for the 1.4 GHz samples (see text).
final two components largely just account for scatter around the
two dominating independent relations between $P$–$z$ and $D$–$\alpha$. This
is an important finding, which is consistent with that demonstrated
by the previous section, that the $P$–$z$ correlation remains dominant
even when a large collection of complete samples is used.

Although PCA should successfully identify all underlying indepen-
dent correlations in the data, we ran the analysis again on just the
subsamples in a well-covered region of the $P$–$z$ plane, thereby re-
moving the selection effect (see Fig. 1). A second motivation for
doing this is to restrict analysis to only high-power radio sources, thus
studying a relatively uniform population (extended double radio
sources), with little contamination from low-power sources which
are often unresolved (see e.g. Baldi & Capetti 2009). In this case,
the observed variance can be attributed to two independent rela-
tions, each giving an almost equal contribution to the variance (see
Table 3, lower). The largest contributor, at 33 per cent, is an anticor-
correlation between $\alpha$ and $D$ as found for the whole sample, followed
by 28 per cent contribution between $D$ and $z$. The third 24 per cent
contribution arises almost solely along the log $P$ axis, uncorrelated
with the other parameters. This confirms our earlier findings that
$D$–$\alpha$ and $D$–$z$ relations are intrinsic to the data set, irrespective
of the presence of Malmquist bias.

The results for a high-frequency-selected collection of samples
(compared of CENSORS and Hercules) show broadly similar res-
ults, albeit with some difference in the detail (see Table 4). For the
entire sample, the results are very similar, again with approximately
half the variance being accounted for by the $P$–$z$ correlation and a
further 30 per cent by a $D$–$\alpha$ correlation. The main difference is that
this latter correlation is weakened somewhat by the third compo-
nent. If we then select a well-covered section of the $P$–$z$ plane (see
Fig. 1), then similar to the low-frequency data, approximately 40 per
cent of the variance is accounted for by a $D$–$z$ anticorrelation, fol-
lowed by a 32 per cent $\alpha$–$D$, $P$ anticorrelation (see Table 4, lower).

A weaker $P$–$\alpha$ correlation accounts for the large remainder of
the variance, which removes the weak $P$–$\alpha$ anticorrelation contribu-
tion of the second component.

From this analysis, we can tentatively conclude that there are
two firm independent relations present in both data sets, between
$D$ and $\alpha$, and between $P$ and $z$. By utilizing the full $P$–$z$ coverage
subsamples, we confirm that the $D$–$\alpha$ correlation is fully inde-
dependent of Malmquist bias. In the subsamples restricted to be
high-power sources with good $P$–$z$ coverage, a strong $D$–$z$ anticor-
correlation is also seen. That the $D$–$\alpha$ correlation is slightly stronger
in the low-frequency data set and the $D$–$z$ anticorrelation stronger
in the high-frequency sample is most likely due to the differing
types of sources which low- and high-frequency-selected samples
collect. Low-frequency samples will primarily be composed of
lobe-dominated sources, suffering little in the way of orientation
bias, and hence a large proportion of large, steep spectrum sources.
High-frequency-selected samples will include many more beamed,
core-dominated and young GPS/Compact Steep Spectrum (CSS)
sources, and less classical lobed-dominated sources.

The correlation between $D$ and $\alpha$ arises due to aging of the radio
sources. As a source gets older, it increases in size and the spec-
trum steepens with age. The physical cause of the anticorrelation
between $D$ and $z$ is subject to more debate. It could arise as the
result of the environment at high redshift, or as a result of sources
at high redshift being more likely to be younger, and hence smaller
(cf. Blundell & Rawlings 1999). It is interesting to note that despite
a Spearman rank test (see Table 2) suggesting the presence of a
relation between $z$ and $\alpha$ almost as strong as that between $D$ and
$\alpha$, the PCA does not clearly identify an independent $z$–$\alpha$ relation
in either the high- or low-frequency-selected samples, suggesting that
the correlation observed may be largely a result of selection effects.

One which may be present is that between radio power and linear
size. It is thought that radio sources follow tracks on the $P$–$D$ plane,
beginning with high power, small sources, and evolving into lower power, larger sources in time (see e.g. Kaiser & Best 2007). Individual low-frequency samples, which are more sensitive to extended radio lobes, show a trend for radio power to increase as linear size decreases, which could arise from the combination of Malmquist bias and the $D-z$ correlation (and indeed, this correlation weakens substantially once Malmquist bias is removed). In a collection of low-frequency-selected samples, this trend in conjunction with any remaining Malmquist bias and the $D-\alpha$ correlation would naturally lead to an extrinsic contribution to the $z-\alpha$ correlation.

### 5 Large Intrinsic Scatter in $\alpha$

Given the independent trends between spectral index, linear size and redshift, identified both visually and by the PCA analysis, an attempt was made to fit an analytical form to the spectral index using linear size, luminosity and redshift. Again, as detailed in Section 3, we use only sources with an observed spectral index steeper than $-0.5$.

Tables 5 and 6 list the best-fitting coefficients for each relation modelled. We began with very simple linear fits, and progressed to fitting planes modelling all four variables. We can see clearly that both the reduced $\chi^2$ and the residual standard deviation decrease, albeit by a small amount, with the inclusion of additional variables in the model for both the high- and low-frequency-selected samples. A plane fit of all four variables gives the best fit, and the smallest deviation in $\alpha$ residuals for both low- and high-frequency-selected samples. The best-fitting model is illustrated in Fig. 4. Although the plane model manages to successfully remove the trends between spectral index and linear size, radio power and redshift, the key finding is that it is unable to predict the observed $\alpha$. The intrinsic scatter in $\alpha$ is much greater than that arising from any physical trends with other observables present in the data sets.

### Table 5

| Model | Sample | $r \chi^2$ | $\sigma$ | $a_1$ | $a_2$ | $a_3$ | $a_4$ |
|-------|--------|-------------|-----------|--------|--------|--------|--------|
| $\alpha = -0.8$ | Whole | 2.47 | 0.15 | – | – | – | – |
| | $P-z$ | 2.53 | 0.14 | – | – | – | – |
| $\alpha = a_1 + a_2$ | Whole | 1.98 | 0.15 | –0.30(0.03) | –0.75(0.01) | – | – |
| | $P-z$ | 1.04 | 0.14 | –0.17(0.06) | –0.80(0.02) | – | – |
| $\alpha = a_1 + a_2$ | Whole | 1.99 | 0.15 | –0.04(0.01) | 0.40(0.14) | – | – |
| | $P-z$ | 1.02 | 0.14 | –0.02(0.02) | –0.41(0.43) | – | – |
| $\alpha = a_1 + a_2 + a_3$ | Whole | 2.12 | 0.15 | –0.05(0.01) | –0.88(0.01) | – | – |
| | $P-z$ | 0.92 | 0.14 | –0.07(0.01) | –0.94(0.01) | – | – |
| $\alpha = a_1 + a_2 + a_3 + a_4$ | Whole | 1.95 | 0.15 | –0.17(0.05) | –0.03(0.01) | –0.09(0.18) | – |
| | $P-z$ | 1.00 | 0.14 | –0.17(0.06) | –0.01(0.02) | –0.59(0.44) | – |
| $\alpha = a_1 + a_2 + a_3$ | Whole | 1.87 | 0.14 | –0.05(0.01) | –0.06(0.01) | 0.49(0.14) | – |
| | $P-z$ | 0.92 | 0.14 | –0.02(0.02) | –0.07(0.01) | –0.44(0.44) | – |
| $\alpha = a_1 + a_2 + a_3 + a_4$ | Whole | 1.83 | 0.14 | –0.36(0.03) | –0.07(0.01) | –0.80(0.01) | – |
| | $P-z$ | 0.88 | 0.13 | –0.25(0.06) | –0.08(0.01) | –0.86(0.02) | – |

### Table 6

| Model | Sample | $r \chi^2$ | $\sigma$ | $a_1$ | $a_2$ | $a_3$ | $a_4$ |
|-------|--------|-------------|-----------|--------|--------|--------|--------|
| $\alpha = -0.8$ | Whole | 4.10 | 0.20 | – | – | – | – |
| | $P-z$ | 3.70 | 0.19 | – | – | – | – |
| $\alpha = a_1 + a_2$ | Whole | 2.75 | 0.20 | –0.21(0.03) | –0.80(0.01) | – | – |
| | $P-z$ | 1.56 | 0.18 | –0.13(0.04) | –0.81(0.01) | – | – |
| $\alpha = a_1 + a_2$ | Whole | 2.81 | 0.20 | –0.01(0.003) | –0.58(0.09) | – | – |
| | $P-z$ | 1.57 | 0.19 | –0.01(0.01) | –0.57(0.20) | – | – |
| $\alpha = a_1 + a_2 + a_3$ | Whole | 3.20 | 0.21 | –0.05(0.01) | –0.95(0.02) | – | – |
| | $P-z$ | 0.81 | 0.18 | –0.09(0.02) | –1.01(0.04) | – | – |
| $\alpha = a_1 + a_2 + a_3 + a_4$ | Whole | 2.74 | 0.20 | –0.24(0.03) | 0.006(0.003) | –0.95(0.10) | – |
| | $P-z$ | 1.56 | 0.18 | –0.13(0.04) | –0.0001(0.01) | –0.81(0.22) | – |
| $\alpha = a_1 + a_2 + a_3$ | Whole | 3.20 | 0.21 | –0.004(0.01) | –0.05(0.01) | –0.85(0.20) | – |
| | $P-z$ | 0.79 | 0.18 | –0.06(0.03) | –0.09(0.03) | 0.58(0.80) | – |
| $\alpha = a_1 + a_2 + a_3 + a_4$ | Whole | 3.14 | 0.21 | –0.18(0.05) | –0.06(0.01) | –0.90(0.02) | – |
| | $P-z$ | 0.75 | 0.17 | –0.41(0.11) | –0.12(0.01) | –0.90(0.02) | – |
| $\alpha = a_1 + a_2 + a_3 + a_4$ | Whole | 3.02 | 0.20 | –0.63(0.1) | 0.08(0.02) | –0.09(0.01) | –2.90(0.38) |
| | $P-z$ | 0.74 | 0.17 | –0.42(0.13) | 0.01(0.04) | –0.12(0.03) | –0.99(0.94) |
some success, but struggle to reproduce $\alpha$ (see e.g. Barai & Wiita 2007).

It is very clear that the correlations between $\alpha$ and size and between luminosity and redshift are weak. The results of this suggest that the use of spectral index alone is unlikely to be efficient in selecting high-redshift radio sources. The equally strong $D-z$ correlation indicates that inclusion of radio size information may increase the efficiency of selection based solely on radio observables.

6 THE ORIGIN OF THE $\alpha$–$z$ CORRELATION IN FLUX-LIMITED SAMPLES

The tendency for observed spectral indices to steepen with redshift has been attributed to a $k$-correction, where as the source spectrum is redshifted, a steeper part of the spectrum is sampled. How much of an effect this is has been a source of much debate in the literature. It has also been suggested that the strength of the $z$–$\alpha$ correlation increases with frequency, as high-frequency parts of the radio spectrum undergo more significant synchrotron losses.

Klamer et al. (2006) find the majority of their USS sample display no curvature, and also cite the well-studied high-redshift source 4C41.17, at $z = 3.8$, as having a straight radio spectrum from 26 MHz to GHz frequencies. They therefore infer that the $k$-correction is irrelevant for high-$z$ USS sources. However, based on this, we cannot simply conclude that the contribution to any $z$–$\alpha$ correlation from the $k$-correction is negligible. In fact, Fig. 5 shows the radio spectra of all currently known $z > 4$ radio galaxies – the majority of which show some evidence of curvature in the observed radio spectrum. Most of the currently known radio galaxies at $z > 4$ display a compact steep radio spectrum, with curvature occurring at low observed frequencies ($\sim 100$ MHz), data which Klamer et al. did not have for their sample. Bornancini et al. (2007) also confirm the presence of curvature at low MHz frequencies for their USS sample.

To quantify the potential effect of the $k$-corrections, we used two samples, CoNFIG and 7CRS. These two samples have the best multifrequency coverage, and hence most accurately determined radio spectra. Rest-frame spectral indices are calculated from fitting a second-order polynomial ($\log S_\nu = a_1 + a_2 \log \nu + a_3 \log^2 \nu$) to the radio spectrum for each source, and measuring the gradient ($\alpha = a_2 + 2a_3 \log \nu$) at the desired frequency; details of this will be presented in Ker et al. (in preparation). A second-order polynomial fit provides a good fit to the radio spectra of the vast majority of sources in each sample.

We then performed a simple linear fit to the observed and rest-frame spectral index measured at three frequencies as a function of $\log (1 + z)$ for both CoNFIG and 7CRS (see Fig. 6). We performed the fit only on sources with an observed spectral index between 1.4 GHz and 151 MHz less than $-0.5$ and with a well-determined radio spectrum. The gradients of these fits then reflect the strength of the $z$–$\alpha$ correlation present (if any). The results we obtain are striking. For 7CRS, we confirm that the gradient of both the observed and the rest-frame $z$–$\alpha$ correlation increases with the frequency at which $\alpha$ is measured, as first reported by Blundell et al. (1999). We also see that the measured $z$–$\alpha$ correlation is approximately twice as steep in the observed frame than in the rest frame (dependent on frequency). It is also worthwhile noting that for 7CRS, contrary to the $z$–$\alpha$ correlation, the $D$–$\alpha$ correlation strengthens in the rest frame.

Similarly, for CoNFIG we see that the observed-frame correlation can be 50 per cent steeper or more at all frequencies than that measured in the rest frame. However, the increase in gradient with frequency is not seen. We suggest that this is because GHz-selected samples pick out very different proportions of various types of radio source, favouring young GPS/CSS, core and beamed sources (much higher orientation bias).

To test this, we ran the fits again, this time excluding all known quasars, and objects classified as compact in CoNFIG, and sources with a size less than 30 kpc in 7CRS. This ensures that the vast majority of sources included in both samples will be lobe dominated and working against the intergalactic medium (IGM), and are not heavily contaminated by beamed sources or are sources so small that they are still propagating through the medium of their host galaxy rather than the IGM. The results are plotted in the bottom
Figure 5. Radio spectra for the nine highest redshift $z > 4$ radio galaxies known. Fluxes were obtained from the NASA Extragalactic Database, at frequencies ranging from 38 MHz to 8 GHz. The flux scale is offset by a small arbitrary amount for each source to allow the shapes of the radio spectra to be compared. TN J0924 ($z = 5.19$), J1639 ($z = 4.88$) and CEN69 ($z = 4.11$) have spectral indices which imply that they should be detected in the VLSS (Cohen et al. 2007) assuming a straight spectrum; however, all three are not. We measure the noise ($\sigma$) at each source position in the VLSS maps, and take the 2$\sigma$ value as the upper limit 74 MHz flux, and assume a 1$\sigma$ error. Note that, interestingly, the two highest redshift known sources are easily detectable at 1.4 GHz in the NVSS, at 71.5 and 21.8 mJy, respectively, but would not be detected in any currently existing 150 or 74 MHz surveys. The vast majority of the spectra are classified as compact steep spectrum, most flattening towards lower frequencies (four of these are potentially peaked). Only one (7C1814, at the lowest redshift) is confirmed as straight over the frequency range from 74 MHz to 5 GHz. This negates the common assumption that high-redshift USS sources display no curvature over a large frequency range. The sources are ordered in $P_{1.4\text{GHz}}$ (calculated with full curvature information), and it can be clearly seen that the most powerful sources are more likely to display significant curvature. It is also worth noting that all but one of these sources (7C1814) are compact and have $\theta < 6$ arcsec.

Figure 6. The upper panel shows the gradient $'\alpha'$ from fitting $\alpha = a \log (1 + z) + b$ to rest-frame and observed spectral indices for CoNFIG and 7CRS. Whilst both samples show a clear decrease in the gradient of the $z$–$\alpha$ correlation when $k$-corrections are applied, only 7CRS shows a marked increase in strength with the frequency at which $\alpha$ is measured. The lower figure shows the same fit again for CoNFIG and 7CRS, but this time with all known quasars and compact objects (classified as compact in CoNFIG, or less than 30 kpc in size in 7CRS) removed. The gradients become very similar at all frequencies and for both samples when only extended sources are considered. These figures clearly show that the $k$-correction can be responsible for up to 50 per cent of the $z$–$\alpha$ gradient observed in flux-limited radio samples.

panel of Fig. 6, and the difference is clear to see. Both CoNFIG and 7CRS now follow very similar relations, both displaying observed gradients which are approximately twice as strong as the rest-frame gradients, but which are now largely independent of the frequency at which $\alpha$ is measured. It is interesting to note that the strength of the gradient for observed $\alpha$ for both samples is very similar to that determined by Ubachukwu, Ugwoke & Ogwo (1996) for a sample of radio galaxies compiled from the 3CRR and WP85 samples, again excluding compact sources.

Our results confirm that once the $k$-correction is removed, a weak correlation between $z$ and $\alpha$ remains for extended radio galaxies, which would fit in with a scenario where lobes are working against a denser environment at higher redshift, and hence high-frequency losses are greater. Miley & De Breuck (2008) note, however, that it is very difficult to reproduce the observed $z$–$\alpha$ relation from this somewhat simplistic density-dependent effect. They suggest that as the density of gas around high-redshift sources has been observed to be highly inhomogeneous, and denser close to the nucleus of the galaxy, the ultrasteep radio spectra are produced by some as yet unknown mechanism within the host galaxy, rather than by the IGM conditions through which the radio lobes are propagating.

We also conclude that GHz-selected samples have a much greater orientation bias present, which can disguise the presence of the $z$–$\alpha$ correlation displayed by extended radio galaxies. We have also successfully demonstrated that the $k$-correction is not negligible when measuring the strength of any $z$–$\alpha$ correlation, and can be responsible for more than 50 per cent of the strength of the observed gradient in a flux-limited sample.
Table 7. Observable parameters for all spectroscopically confirmed radio galaxies at redshift $z > 3$ in all of the samples studied. A CSS radio spectrum indicates that the source is compact, steep and peaks at low frequencies. A C- spectrum displays negative curvature, but no peak within the observed frequency range.

| Name     | Sample | $z$  | $K$  | $\alpha$ | D/kpc | Radio spectrum |
|----------|--------|------|------|----------|-------|----------------|
| 7C1745+6624 | 7CRS   | 3.01 | 20.25| -0.78    | 3.85  | CSS            |
| TOOT0–1214  | TOOT00 | 3.081| 18.6 | -1.13    | 115   | C-             |
| CEN 16      | CENSORS| 3.126| 19.32| -0.86    | 99.6  | C-             |
| 7C1748+6703 | 7CRS   | 3.2  | 18.27| -0.97    | 106   | C-             |
| 6C1232+3942 | 6CE    | 3.22 | 17.82| -1.14    | 228   | C-             |
| CEN 105     | CENSORS| 3.38 | 20.16| -1.16    | 50    | Straight       |
| 6C 0902+3419| 6CE    | 3.4  | 19.70| -0.84    | 91.3  | Straight       |
| CEN 24      | CENSORS| 3.43 | 19.30| -0.66    | 10    | CSS            |
| 7C1814+6702 | 7CRS   | 4.05 | 19.16| -1.01    | 124   | Straight       |
| CEN 69      | CENSORS| 4.11 | 19.60| -0.80    | 7.2   | C-             |

7 IMPLICATIONS FOR HIGH-REDSHIFT SEARCHES

The data collected for the nine complete samples allow us to measure for the first time the efficiency of the three most commonly used methods in the literature for searching for high-redshift radio galaxies, namely radio spectral index, angular size and $K$-band magnitude. We are looking for a set of criteria which minimizes the size of the selected subsamples that would require follow-up observations, whilst maximizing the number of high-$z$ galaxies retained in this sample. We assume a definition of highest efficiency as maximizing the difference between the total fraction of the sample recovered and the total fraction of high-$z$ sources recovered, with each increasing cut in the selection parameter under study. We choose to consider $z > 2$ radio galaxies as high-$z$ sources, as for the data sets under consideration this provides the optimal compromise between maximizing the redshift whilst still maintaining sufficient high-$z$ sources to allow a robust analysis. For comparison, we also show the analysis repeated for $z > 3$ where possible, albeit with much lower number statistics [we have 10 $z > 3$ radio galaxies with spectroscopic redshifts (see Table 7) and six with photometric redshifts in our samples]. As there are only approximately 50 radio galaxies with $z > 3$ known (Ishwara-Chandra et al. 2010), our samples are hence representative of the highest known redshift radio galaxy parameter space.

7.1 Spectral index selection

As discussed above, an initial steep radio spectral index cut is an extremely popular method of reducing very large radio samples down to manageable sizes for imaging and spectroscopic follow-up, in order to locate high-redshift sources. We now investigate whether a first spectral index cut does indeed recover a significant proportion of high-$z$ sources present in the samples. Many recent studies in the literature base searches for high-$z$ radio galaxies on the assumption that they may be distinguished by a steep spectrum. Ishwara-Chandra et al. (2010) provide a list of the highest redshift, $z > 3$, known radio galaxies, 47 in number, the vast majority of which have been selected from a USS sample. However, Jarvis et al. (2009) also recently reported the discovery of the second highest redshift radio galaxy known, a source which they noted clearly does not have an ultrasteep spectral index (see Fig. 5). Work with the DRaGONS study (Schmidt, Connolly & Hopkins 2006), which uses a large, bright radio sample from the 1.4 GHz Faint Images of the Radio Sky at Twenty Centimetres (FIRST) survey with redshifts estimated from the $K$–$z$ relation, also suggests that even with a relatively flat spectral index selection criterion of $\alpha < -0.8$, one-third of the $z > 2$ sources are missed. In Table 7, we present a list of the 10 radio galaxies with a confirmed spectroscopic redshift of $z > 3$ from all of the samples used in this study. Five of these do have a steep $\alpha < -1.0$; however, the remainder display a wide variety of spectral indices.

In studies utilizing an ultrasteep selection criterion, an often used argument to justify the use of steep spectral index cut-offs is the apparent strong shift in redshift distribution to high redshift. However, in the majority of these studies (e.g. De Breuck et al. 2000; Bryant et al. 2009), the samples are very large (numbering in the hundreds), and spectroscopic follow-up is expensive, and so often faint $K$- or $11\mu m$-band detections or limits are used as additional selection criteria when deciding which targets to pursue with spectroscopy. This makes it very difficult to disentangle the extent to which the USS or the optical/near-infrared selection criterion is responsible for preferentially selecting high-redshift sources.

Armed with spectroscopically complete samples at a variety of flux density limits and finding frequencies, we can determine the efficiency of the USS selection technique in an unbiased way, for an observed $\alpha \geq -1.2$ (we have too few sources steeper than this to study robustly). In Fig. 7, we take each sample in its entirety, blindly apply a decreasing spectral index limit, and calculate the median redshift of the resulting sample of sources steeper than that limit. Considering first the results with no cuts applied, Fig. 7 offers a clear observational confirmation that the redshift distributions of complete samples are dependent on the corresponding flux density limits of each sample and selection frequency. It is immediately apparent that low-frequency-selected samples, even at relatively bright flux density limits, select on average higher redshift sources. For both low- and high-frequency-selected samples, applying a cut of $\alpha < 0.9$ generally increases the median redshift of the obtained sample. That the median redshift decreases at very steep spectral index cuts is most likely due to the inclusion of very steep spectrum, low-redshift clusters (e.g. De Breuck et al. 2000). The samples selected at increasingly faint flux density limits for both high- and low-frequency-selected samples also display higher median redshifts, except for the faint Hercules sample at high frequency. As described in Best et al. (2003), the CENSORS flux limit was chosen because, according to the models of Dunlop & Peacock (1990), a survey with a flux density limit of approximately 10 mJy at 1.4 GHz is optimal for detecting sources at redshifts greater than 2.5, with the percentage of high-$z$ sources detected decreasing at lower and higher flux densities. Our results for the GHz-selected samples are consistent with this. Most recently, Afonzo et al. (2011) have extended the search for USS, high-redshift sources to the sub-mJy level, and their sample appears to be broadly consistent in expected content with the much brighter samples studied here. Also illustrated clearly by Fig. 7 is the fact that the median values of redshift obtained from existing USS-selected samples are very similar to those we obtain for USS selection of the complete samples used in this study [excepting the De Breuck et al. (2000) USS sample, which includes a very strong additional selection criteria of targeting only those sources with the faintest $K$-band magnitudes].

In Fig. 8, we consider the efficiency and completeness of a USS criterion in selecting high-$z$ radio galaxies. The figure shows the proportion of high-$z$ sources recovered in each sample as steeper spectral index cuts are applied, along with the proportion of the entire sample that is returned, and the overall high-$z$ content of the...
Figure 7. The median redshift of all sources in a given sample which have spectral index steeper than a given spectral index, as a function of the spectral index limit. Data points are only plotted if the remaining sample size is at least five sources. The left-hand panel displays the 1.4 GHz selected samples, and the right-hand panel shows the 150 MHz selected samples. Highlighted in stars are the De Breuck et al. (2000) 1.4 GHz USS sample median redshift and the median redshifts of the MRCR-SUMSS (Bryant et al. 2009), NVSS-SUMSS (Klamer et al. 2006), 4C USS (Chambers et al. 1996) 6C* (Jarvis et al. 2001) and 6C** (Cruz et al. 2007) USS samples. These samples have additional biases due to $K$-band selection and incomplete spectroscopic redshifts, and hence a direct comparison is not possible, but it is interesting to note that their median redshifts are broadly consistent with what we see in spectral index cut, spectroscopically complete samples.

Figure 8. The bottom panel shows the fraction of all sources (solid lines) and the fraction of $z > 2$ (dashed) and $z > 3$ (dotted) radio galaxies that have steeper spectral indices than the given limit, as a function of that limit, for both the high- (black) and low-frequency-selected (red) samples. The top panel displays the fraction of high-$z$ radio galaxies in the sample recovered by these cuts. Note that the $z > 3$ lines are much more uncertain, due to the relatively low numbers of these (see text).

Reduced subsample. This analysis is carried out using 7CRS and TOOT00 samples at low frequency, and CENSORS and Hercules at high frequency. The brighter samples are not included as they have very few $z > 2$ sources, and the combination of CENSORS and Hercules, 7CRS and TOOTS00 provides two large samples of approximately 150 sources each. The results of this are very interesting: for the low-frequency-selected sample, the baseline 15 per cent fraction of high-$z$ sources in the recovered subsample nearly doubles to 30 per cent with a spectral index cut $\alpha = -1$, but at a cost of removing 60–70 per cent of the known high-$z$ sources from the recovered subsample. For the high-frequency-selected samples, there is hardly any difference, regardless of the spectral index cut applied. In other words, for the high-frequency-selected sample, by excluding sources flatter than the cut, we are not gaining a substantial proportion of high-$z$ sources above that which we would expect if there was no correlation, and the data were distributed evenly across the $\alpha-z$ plane.

In utilizing complete samples to address the question of high-$z$ selection efficiency using USS samples, our main limitation is the low number of extreme spectral index sources included in our collection of complete samples. Any radio sample will include $\sim 5$ per cent $\alpha < -1$ and $\sim 1$ per cent $\alpha < -1.3$ (e.g. De Breuck et al. 2000), and indeed these proportions hold for the samples presented here: we have too few sources to study the $\alpha < -1.2$ range. What is needed is complete spectroscopic follow-up of USS samples encompassing these extreme spectra sources, in order to determine the high-$z$ fraction. However, very few USS samples available in the literature have substantial spectroscopic completeness. Most have additional optical or angular size biases applied (e.g. Röttgering et al. 1996; De Breuck et al. 2000) when selecting candidates for spectroscopy. Chambers et al. (1996) present one of the most complete USS samples available. They study a small sample of 4C USS sources ($\alpha < -1.0$), selected from Tielens et al. (1979) which is 50 per cent spectroscopically complete, with 15 having $R$- or $I$-band magnitudes, one with an $I$-band limit and one with no magnitude data. There are eight sources with spectroscopic redshift $z > 2$. From the magnitude distribution, it is likely that those without spectroscopic redshifts are in the range $1.0 < z < 1.6$. This gives the fraction of the USS sample with $z > 2 = 24$ per cent which is very much in line with our findings for low-frequency samples with this spectral index limit (see Figs 7 and 8).
We can conclude from this that a USS selection criterion does work at low frequency, but is not a strong effect, whilst it is inefficient for high-frequency-selected samples.

7.2 Angular size selection

Another often used criteria for maximizing the high-z content of radio source samples is that of angular size. In Fig. 9, we plot a similar diagram to that of the spectral index cuts. In this, it is clear that moderate cuts can be made to the sample based on angular size, whilst still ensuring the large majority of high-redshift sources remain.

The fraction of high-z sources in the recovered sample is similar for both samples, remaining constant at ~15 per cent for the majority of angular size cuts, and increasing up to ~25 per cent for angular size cuts less than 10 arcsec. Contrary to a spectral index cut, angular size cuts prove to be generally more effective for high-frequency-selected samples. For example, applying a cut of 5 arcsec retains 70 per cent of known high-z sources in the sample, whilst reducing the total sample to 40 per cent of its original size. However, a similar cut for the low-frequency sample retains less than 40 per cent of the known high-z sources in the sample. Despite the high-z fraction in the remaining subsample having nearly the same dependence on the θ cut for both high- and low-frequency-selected samples, at low frequency a much smaller proportion of the total high-z galaxies is recovered.

We conclude that angular size cuts can successfully retain the majority of high-z sources, whilst almost halving the original sample size for high-frequency-selected samples. A larger angular size cut must be applied to low-frequency samples in order to retain the same efficiency as seen for high-frequency samples with a smaller cut applied. However, once again this is not a particularly efficient technique.

7.3 K-band selection

Selecting high-redshift galaxies from near-infrared imaging is possible, thanks to the very tight relation observed between K-band magnitudes and redshift (e.g. Lilly & Longair 1982; Willott et al. 2003). It is worth noting too here that developments in recent years have identified a new population of radio sources without optical or infrared detections, infrared-faint radio sources (IFRS). These are potentially excellent very high-redshift candidates (see e.g. Middelberg et al. 2011; Garn & Alexander 2008). The second highest redshift known radio galaxy identified by Jarvis et al. (2009) is an IFRS, and was selected for follow-up based purely on its lack of optical or K-band detection (it has a spectral index which is not USS, α = −0.75).

The main drawback with this method, however, is that very deep K-band imaging is required over the radio survey area. To illustrate the efficiency of K-band imaging in selecting high-redshift radio sources, we carry out a similar analysis to that performed for spectral index and angular size. We utilize only CENSORS and 7CRS for this analysis, as both samples are highly spectroscopically complete, and have readily available K-band data. For both of these samples, in addition to the −0.5 spectral index cut as detailed in Section 4, we also exclude all known radio quasars in both samples, as these do not follow the K–z relation of radio galaxies. The high spectroscopic completeness is necessary, as we only want to use sources with spectroscopic redshifts, and not those with redshifts estimated from the K–z relation. The aperture-corrected K-band data for CENSORS are taken from Brookes et al. (2006) and Rigby et al. (2011), and for 7CRS, aperture-corrected K-band data were obtained from publicly available online catalogues.

It should be noted that 7CRS does not have complete K-band data for the sample, with 26 of the 92 radio galaxies having no K magnitude. However, all but one of these sources without a K measurement are at redshift 1 or below, and given the very tight K–z relation, all of these are expected to be bright, K < 17 sources, and should not significantly affect the analysis of high-z sources in this sample.

In Fig. 10, we plot the fraction of high-z sources recovered with an increasing K-band magnitude cut for CENSORS and 7CRS. It is immediately clear that a cut of 18.5 in K-band magnitude recovers almost all high-z sources for both samples, with very few low-redshift sources included. In previous years, applying this technique required dedicated deep K-band surveys, expensive in telescope time: cross-matching with existing wide-area K-band surveys such as the 2 Micron All Sky Survey (2MASS) would potentially reduce the sample size by 10–20 per cent, but this is limited by the bright K-band magnitude limits. The release of the UKIRT Infrared Deep Sky Survey (UKIDSS) (Lawrence et al. 2007) Large Area Survey data mitigates this somewhat, as the K limit reaches 18th magnitude (Vega), and covers many thousands of square degrees in sky area.

If we apply an 18th magnitude limit to our samples, then all the high-z sources are recovered, whilst the sample is reduced to ~30 per cent of its original size. This is a far more successful selection method than any based on radio properties alone, and is now feasible over large sky areas and high resolution, wide and deep radio surveys (limited to GHz frequencies) to enable matching of host galaxy to

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Figure 9. The bottom panel shows the fraction of all sources (solid lines) and the fraction of z > 2 (dashed) and z > 3 (dotted) radio galaxies that have smaller angular sizes than the given limit, as a function of that limit, for both the high- (black) and low-frequency-selected (red) samples. The top panel displays the fraction of high-z radio galaxies in the sample recovered by these cuts. Note that the z > 3 lines are much more uncertain, due to the relatively low numbers of these.

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4 https://www.astrosci.ca/users/willottc/kz/kz.html
uncertain, due to the relatively low numbers of these. Out using a reduction of the numbers of these cuts. Note that the magnitude alone. Plotted in red stars are the highest redshift radio galaxies known with K, z as described in the text. It is immediately clear that the combination of both radio observables is much more effective in predicting the source redshift. The bottom panel shows the fraction of all radio galaxies (solid lines), and the fraction of z > 2 (dashed), z > 3 (dotted) radio galaxies that have fainter K-band magnitudes than the given limit (bins of 1 mag), as a function of that limit, for both CENSORS (black) and 7CRS (red) samples. The top panel displays the fraction of z > 2, z > 3 radio galaxies in the two samples recovered by these cuts. Note that the z > 3 lines are much more uncertain, due to the relatively low numbers of these.

7.4 Optimal search criteria for high-z radio galaxies

Many combinations of cuts using the K-band magnitude, angular size and spectral index have been utilized in the literature, but as yet there have been no investigations into the most efficient combination of these for selecting high-z radio galaxies. As we have shown previously, there are some correlations present between D, α and z in flux-limited samples, in addition to the well-known K-z relation.

We therefore test whether fitting a simple relation to these observed parameters would enable a more efficient selection to be made. We fitted a function first to angular size and observed spectral index (i.e. a radio-only selection method), and then to angular size, spectral index and K-band magnitude (just for the radio galaxies, cf. previous section) as follows:

\[ \log(1+z) = a_1 \log \theta + a_2 \alpha + a_3 \]  

(1)

\[ \log(1+z) = a_1 (K-18)^2 + a_2 (K-18) + a_3 \log \theta + a_4 \alpha + a_5 \]  

(2)

Having obtained best-fitting parameters, we used these two relations to derive a predicted redshift, \( z_{\text{pred}} \), for each source (see Table 8).

Table 8. The fits to \( z(\theta, \alpha) \) and \( z(K, \alpha) \) for 7CRS and CENSORS as described in the text.

| Sample | Function | \( \mu(z_{\text{obs}} - z_p) \) | \( \sigma(z_{\text{obs}} - z_p) \) | a1 | a2 | a3 | a4 | a5 |
|--------|----------|----------------------|----------------------|-----|-----|-----|-----|-----|
| 7CRS   | \( \log(1+z) = a_1 \log \theta + a_2 \alpha + a_3 \) | -0.26 | 0.64 | -0.1213 | -0.5840 | 0.045 | - | - |
| CENSORS| \( \log(1+z) = a_1 \log \theta + a_2 \alpha + a_3 \) | -0.45 | 0.88 | -0.1687 | -0.1973 | 0.40725 | - | - |
| CENSORS| \( \log(1+z) = a_1 (K-18)^2 + a_2 (K-18) + a_3 \log \theta + a_4 \alpha + a_5 \) | -0.08 | 0.40 | 0.011 | 0.11287 | -0.0497 | 0.0024 | 0.4074 |
| 7CRS   | \( \log(1+z) = a_1 (K-18)^2 + a_2 (K-18) + a_3 \log \theta + a_4 \alpha + a_5 \) | -0.12 | 0.47 | 0.004 | 0.1025 | -0.003 | -0.1327 | 0.2963 |

Figure 11. Each panel plots the predicted redshift versus the actual redshift for CENSORS (dots) and 7CRS (squares), for three different relations. The left-hand panel plots the predicted redshift from a simple linear fit to spectral index, and the middle the predicted redshift from fitting z as a function of \( \theta \) and \( \alpha \) as described in the text. It is immediately clear that the combination of both radio observables is much more effective in predicting the source redshift. The panel on the right displays the predicted redshift from fitting z as a function of \( K, \theta \) and \( \alpha \), a relation which offers little improvement over a simple fit to \( K \) magnitude alone. Plotted in red stars are the highest redshift radio galaxies known with K, \( \theta \) and \( \alpha \) data available, all of which would be successfully singled out using a \( z_{\text{pred}}(K, \theta, \alpha) \).
The results of this can be seen in Fig. 11. The combination of radio observables does far better than fitting only one single radio parameter (spectral index) alone, whereas, in contrast, the addition of radio variables to the K-band function provides little discernable improvement over fitting K-band magnitudes alone. Applying these findings, we then repeated the analysis of the previous subsections by applying increasing predicted redshift cuts from these two relations. The results are shown in Figs 12 and 13. Whilst the $z(\theta, \alpha)$ relation does not give a perfect fit to the data (see Table 8), applying cuts based on the predicted redshifts results in a substantially higher efficiency than any one radio variable cut alone (see Fig. 12). The $z(\theta, \alpha)$ fit is less efficient for the high-frequency sample CENSORS than for the low-frequency-selected 7CRS, as we would expect from the findings of the preceding subsections. The $z(K, \theta, \alpha)$ fit appears equally efficient for both (note that 7CRS falls off more quickly in Fig. 13, as it contains far less sources above $z = 3$ than CENSORS), but on comparison with a simple K magnitude fit (see Fig. 10), any improvement is very marginal.

As a final test, we also calculate $z(\theta, \alpha)$ and $z(K, \theta, \alpha)$ for the nine highest redshift radio sources known (see Fig. 11). For all of these sources, the $z(K, \theta, \alpha)$ and $z(\theta, \alpha)$ relations predict high redshifts, $z > 2$, which, if we applied as cuts to a complete sample of radio galaxies, would leave only a very small proportion of the original sample.

**8 CONCLUSIONS**

The main conclusions of this paper are as follows.

(i) The strongest independent relation measured in both high- and low-frequency-selected samples, excluding the $P-z$ correlation (which is a selection effect), is between $D$ and $\alpha$

(ii) The observed $z-\alpha$ correlation reaches maximum strength for an observed $\alpha$ measured at high frequencies, in a low-frequency-selected sample. However, this correlation is weak in comparison to the other observed correlations between $\alpha-D$ and $D-z$.

(iii) Up to 50 per cent of the measured $z-\alpha$ gradient can be contributed by a $k$-correction, in both high- and low-frequency-selected samples. This is important as almost all known $z > 4$ galaxies display curvature in their spectrum.

(iv) Selecting high-redshift ($z > 2$) sources based only on their observed $\alpha$ provides only a small increase in searching efficiency for low-frequency-selected samples, and almost none for high-frequency-selected samples. Table 9 displays the fraction of the samples that have $z > 2$ and 3 for a selection of the observational cuts studied.

(v) Whilst we confirm the presence of a $z-\alpha$ correlation for extended classical radio galaxies, if it arises as a result of radio lobes working against an increasingly denser IGM, giving a steeper spectrum, we caution that this may not be as useful at the very highest redshifts. The very highest $z > 4$ known radio sources present observational characteristics which are more consistent with being young radio sources, still confined within their host galaxies.

(vi) K-band selection is very much more efficient than radio-based selection to maximize the number of high-$z$ galaxies selected whilst minimizing the total sample size. Recent existing surveys such as UKIDSS Large Area Survey are just deep enough to enable efficient searches.

(vii) Searching based on a combination of criteria such as near-infrared magnitude, radio spectrum and $\theta$ provides optimal searching efficiency for all types of radio source at high redshift.

The key finding of this paper is that the efficiency of the USS criterion alone in selecting the highest redshift radio galaxies is not
Table 9. The fraction of high-$z$ sources in samples with various observational cuts applied. Column 1 displays the samples used, the second and third columns display the fractions of $z > 2$ and 3 radio galaxies in the whole sample, and the fourth column the observational cut (in spectral index, size, K-band or predicted redshift from a combination of these) applied to the sample(s) used. Column 5 displays the total fraction of the sample(s) that is returned by applying the observational cut, and the final two columns display the fractions of $z > 2$ and 3 radio galaxies in the returned sample.

| Samples used | Per cent of whole sample at $z > 2$ (per cent) | Per cent of whole sample at $z > 3$ (per cent) | Observational cut applied | Per cent of retained sample at $z > 2$ (per cent) | Per cent of retained sample at $z > 3$ (per cent) | Per cent of sources at $z > 2$ lost by cut (per cent) | Per cent of sources at $z > 3$ lost by cut (per cent) |
|---------------|-----------------------------------------------|-----------------------------------------------|---------------------------|-----------------------------------------------|-----------------------------------------------|-----------------------------------------------|-------------------------------------------------|
| CEN, Her      | 13                                            | 4                                             | $\alpha < -1.0$           | 10                                           | 20                                           | 8                                            | 62                                             |
| 7C, TOOT      | 12                                            | 4                                             | $\alpha < -1.0$           | 12                                           | 20                                           | 6                                            | 68                                             |
| 7C, TOOT      | 13                                            | 4                                             | $\theta < 10$            | 13                                           | 55                                           | 6                                            | 19                                             |
| CEN           | 15                                            | 5                                             | $\alpha < -1.0$           | 30                                           | 20                                           | 4                                            | 47                                             |
| 7C            | 11                                            | 3                                             | $\theta < 10$            | 19                                           | 30                                           | 13                                           | 47                                             |
| 7C            | 18                                            | 6                                             | $K > 19$                  | 15                                           | 35                                           | 12                                           | 54                                             |
| 7C            | 21                                            | 5                                             | $K > 19$                  | 10                                           | 83                                           | 30                                           | 8                                              |
| CEN           | 18                                            | 6                                             | $\alpha < -1.0$           | 20                                           | 80                                           | 28                                           | 62                                             |
| 7C            | 21                                            | 5                                             | $\alpha < -1.0$           | 20                                           | 62                                           | 16                                           | 38                                             |

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New insights on the z–α correlation

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