CONVEXIFICATION WITH BOUNDED GAP FOR RANDOMLY PROJECTED QUADRATIC OPTIMIZATION

TERUNARI FUJI, PIERRE-LOUIS POIRION†, AND AKIKO TAKEDA§

Abstract. Random projection techniques based on Johnson-Lindenstrauss lemma are used for randomly aggregating the constraints or variables of optimization problems while approximately preserving their optimal values, that leads to smaller-scale optimization problems. D’Ambrosio et al. have applied random projection to a quadratic optimization problem so as to decrease the number of decision variables. Although the problem size becomes smaller, the projected problem will also almost surely be non-convex if the original problem is non-convex, and hence will be hard to solve.

In this paper, by focusing on the fact that the level of the non-convexity of a non-convex quadratic optimization problem can be alleviated by random projection, we find an approximate global optimal value of the problem by attributing it to a convex problem with smaller size. To the best of our knowledge, our paper is the first to use random projection for convexification of non-convex optimization problems. We evaluate the approximation error between optimum values of a non-convex optimization problem and its convexified randomly projected problem.

Key words. random projection, Johnson-Lindenstrauss lemma, quadratic programming, non-convex optimization

AMS subject classifications. 90C20, 90C26

1. Introduction. We consider the following non-convex quadratic optimization problem having large-scale decision variables $x \in \mathbb{R}^n$:

\begin{equation}
\mathbf{P} \equiv \min_x \{ x^T Q x + c^T x \mid A x \leq b \},
\end{equation}

where $Q \in \mathbb{R}^{n \times n}$ is a symmetric matrix, $A \in \mathbb{R}^{m \times n}$, $c \in \mathbb{R}^n$ and $b \in \mathbb{R}^m$. Here, we assume that the feasible region is full dimensional (hence, equality constraints are excluded) and at least one of the eigenvalues of $Q$ is positive.

In this paper, we find a feasible solution with a bounded approximation error to the optimal value of (1.1). For the purpose, we reduce the non-convex problem to a lower-dimensional convex optimization problem using random projection and convexification techniques, and evaluate the gap between optimum values of the two optimization problems. Random projection refers to the technique that maps a set of points $X \subseteq \mathbb{R}^n$ to a set $PX \subseteq \mathbb{R}^d$ in a lower dimensional subspace with random matrices $P \in \mathbb{R}^{d \times n}$ in a way that some intrinsic properties of the set $X$ are approximately preserved with high probability. The main idea of random projections comes from the Johnson-Lindenstrauss lemma ([9]) that states that if the probability distribution of $P$ is properly chosen then there exists $d < n$ such that the Euclidean distance between any pair $x, y \in X$ of points in $X$ is approximately preserved with high probability, i.e. $\|Px - Py\| \approx \|x - y\|$.

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†Graduate School of Information Science and Technology, The University of Tokyo, Tokyo 113-8656, Japan. (terun0818258@g.ecc.u-tokyo.ac.jp).

‡Center for Advanced Intelligence Project, RIKEN, Tokyo 103-0027, Japan. (pierre-louis.poirion@riken.jp).

§Graduate School of Information Science and Technology, The University of Tokyo, Tokyo 113-8656, Japan and Center for Advanced Intelligence Project, RIKEN, Tokyo 103-0027, Japan. (takeda@mist.i.u-tokyo.ac.jp).
Random projections have been already used in various studies to reduce the size of an input matrix while retaining most of its information (see for example [25, 7, 6]), and they are often used for some machine learning problems. Notice that this framework may also be referred as sketching, if random projection matrices are used not to reduce the dimension of the decision variables $x$ but to reduce the sample size of the data in the problem. For example in a least-square problem setting, i.e. $\min_{x \in C} \|y - Xx\|^2$ where $C \subseteq \mathbb{R}^n$ is a convex set, $X \in \mathbb{R}^{m \times n}$ is the design matrix and $y \in \mathbb{R}^{m'}$ is the response vector, the sample size $m'$ is reduced using random projections. The use of random projections to approximate least-squares problems has been extensively studied by, for example, [15, 3, 18, 24]. Random projections have also been used in a non-convex setting: in [2], the authors apply random projections for the $k$-means clustering problem to reduce the number of data points, i.e., the sample size in the problem. Since the reduced problem is also a non-convex optimization problem, the error from its optimum value is evaluated under the assumption that an approximation algorithm is used.

While random projection has also been applied to reduce the number of constraints of a Linear Problem (LP) by randomly aggregating them in [23], it has been applied to a Quadratic Problem (QP) so as to decrease the number of decision variables in [5]. In [1] the authors apply random projections to Semi-Definite Programming (SDP): the variables of the SDP are randomly projected to a space of lower dimension.

In this paper, we show that random projections can also be applied to convexify a non-convex optimization problem. More precisely, we will use random projection to define a convexification of (1.2) and give some error bounds for the error between these two problems. Notice that in [5] the authors already use a random projection matrix $P \in \mathbb{R}^{d \times n}$ to project (1.1) into the following QP:

$$\text{RP} \equiv \min_{u} \left\{ u^T \bar{Q} u + \bar{c}^T u \mid \bar{A} u \leq b \right\},$$

where $u \in \mathbb{R}^d$, $\bar{Q} = P Q P^T$, $\bar{c} = P c$ and $\bar{A} = A P^T$. However, although (1.2) is a QP of smaller size, i.e. the variables of (1.2) belong to a smaller dimensional space, if (1.1) is non-convex then the projected problem will also almost surely be non-convex, and hence will be hard to solve. Therefore, we focus on the fact that if $d$ is small enough then eigenvalues of the matrix $\bar{Q}$ are skewed towards positive values (see Figure 1 shown later), which implies that ignoring negative eigenvalues for the reduced matrix due to the convexification does not lose much information about problem (1.1). In this paper, taking advantage of the fact, we show the following: if the dimension $d$ is carefully chosen then (1.1) can be approximated by a convex QP of smaller size. More precisely, we consider the following convex QP:

$$\text{CRP} \equiv \min_{u} \left\{ u^T \bar{Q}^+ u + \bar{c}^T u \mid \bar{A} u \leq b \right\},$$

where $\bar{Q}^+ = \mathcal{F}^+(\bar{Q})$ is the projection of $\bar{Q}$ onto the positive semidefinite cone. Using an optimum $u^*$ of (1.3), we have a feasible solution $P^T u^*$ to (1.1), for which an approximation error from the optimum value of (1.1) is estimated.

To the best of our knowledge, our paper is the first to use random projection for convexification of non-convex optimization problems. We evaluate the approximation error between optimum values of a non-convex optimization problem and its convexified problem. More precisely, we will prove that if the dimension $d$ is properly chosen then the optimal value of $\text{CRP}$ is a good approximation of the one of $P$. 
The rest of this paper is organized as follows. In section 2, we introduce mathematical preliminary. In section 3, we prove our main results on approximate optimality under the assumption that \( \text{tr} \, Q > 0 \) and in section 4, we discuss how to relax the assumption while achieving similar theoretical results. In section 5 we discuss the results of numerical experiments for two types of problems: randomly generated problems and support vector machine (SVM) problems with indefinite kernel which are attributed to non-convex quadratic optimization problems. Conclusions follow in section 6.

All the notations used in this paper are in Table 1.

| Notation | Convention |
|----------|------------|
| \( C_0, C_1, C_2, C_3 \) | absolute constant |
| \( C \) | \( C = \max\{C_2, C_3\} \) |
| \( \|X\|_{\psi_2} \) | the sub-Gaussian norm of a sub-Gaussian random variable |
| \( \|X\|_{\psi_1} \) | the sub-exponential norm of a sub-exponential random variable |
| \( |a| \) | the Euclidean norm of a vector \( a \) |
| \( \|M\| \) | the operator norm of a matrix \( M \): \( \|M\| = \max_{\|x\|=1} \|Mx\| \) |
| \( \|M\|_F \) | the Frobenius norm of a matrix \( M \): \( \|M\|_F = \sum_{ij} M_{ij}^2 \) |
| \( F^+(M) \) | the projection onto the positive semidefinite cone of a matrix \( M \) |
| \( \mathbf{1} \) | the all one vector |
| \( I_n \) | the identity matrix of size \( n \) |
| \( \text{diag}(a) \) | the matrix whose diagonal is the vector \( a \) |
| \( \text{cond}(M) \) | the condition number of a matrix \( M \) |
| \( \text{opt}(F) \) | the optimal value of an optimization problem \( F \) |
| \( \mathbb{E}(X) \) | expectation of a random variable \( X \) |
| \( \delta_{ij} \) | Kronecker delta: \( \delta_{ij} = 1 \) if \( i = j \), \( \delta_{ij} = 0 \) otherwise |
| \( \mathcal{N}(\mu, \Sigma) \) | the normal distribution with mean \( \mu \) and covariance \( \Sigma \) |

2. Preliminaries.

2.1. Sub-Gaussian and sub-exponential random variables. In this section we review some necessary definitions and theorems in the paper. First, we recall some properties of sub-Gaussian and sub-exponential random variables and concentration inequalities.

**Definition 2.1** (Sub-Gaussian random variables). A random variable \( X \) that satisfies

\[
\mathbb{E}[\exp(X^2/K^2)] \leq 2
\]

for some \( K > 0 \) is called a sub-Gaussian random variable. The sub-Gaussian norm of \( X \), denoted \( \|X\|_{\psi_2} \), is defined to be the smallest \( K \) that satisfies the above inequality, or equivalently, we define

\[
\|X\|_{\psi_2} = \inf\{s > 0 \mid \mathbb{E}[\exp(X^2/s^2)] \leq 2\}.
\]

**Lemma 2.2** ([22, Example 2.5.8]). A Gaussian random variable \( X \sim \mathcal{N}(0, \sigma^2) \) is sub-Gaussian with \( \|X\|_{\psi_2} \leq C_2 \sigma \), where \( C_2 \) is an absolute constant.
DEFINITION 2.3 (Sub-exponential random variables). A random variable $X$ that satisfies
\[ \mathbb{E}[\exp(|X|/K)] \leq 2 \]
for some $K > 0$ is called a sub-exponential random variable. The sub-exponential norm of $X$, denoted $\|X\|_{\psi_1}$, is defined to be the smallest $K$ that satisfies the above inequality, or equivalently, we define
\[ \|X\|_{\psi_1} = \inf\{s > 0 \mid \mathbb{E}[\exp(|X|/s)] \leq 2\}. \]

LEMMA 2.4 ([22, Exercise 2.7.10]). For a sub-exponential random variable $Z$, the difference $Z - \mathbb{E}[Z]$ is sub-exponential too, and
\[ \|Z - \mathbb{E}[Z]\|_{\psi_1} \leq C_3 \|Z\|_{\psi_1}, \]
where $C_3$ is an absolute constant.

In the following, we often use the absolute constant $C$ defined by $C = \max(C_2, C_3)$.

Sub-Gaussian and sub-exponential distributions are closely related as we can see in the following lemma, which implies that the product of sub-Gaussian random variables is sub-exponential.

LEMMA 2.5 ([22, Lemma 2.7.7]). Let $X$ and $Y$ be sub-Gaussian random variables, then $XY$ is sub-exponential. Moreover,
\[ \|XY\|_{\psi_1} \leq \|X\|_{\psi_2} \|Y\|_{\psi_2}. \]

We also recall Bernstein’s inequality for sub-exponential random variables.

THEOREM 2.6 (Bernstein’s inequality, [22, Theorem 2.8.1]). Let $X_1, X_2, \ldots, X_N$ be independent, mean zero, sub-exponential random variables. Then, for every $t \geq 0$, we have
\[ \operatorname{Prob}\left( \left| \sum_{i=1}^{N} X_i \right| \geq t \right) \leq 2 \exp\left( -C_1 \min\left( \frac{t^2}{\sum_{i=1}^{N} \|X_i\|_{\psi_1}^2}, \frac{t}{\max_i \|X_i\|_{\psi_1}} \right) \right), \]
where $C_1$ is an absolute constant.

2.2. Definitions of $\varepsilon$-net and estimation of the operator norm of a matrix. Next, we recall the definition of $\varepsilon$-net.

DEFINITION 2.7. Consider a subset $K \subset \mathbb{R}^n$ and let $\varepsilon > 0$. A subset $N \subset K$ is called an $\varepsilon$-net of $K$ if every point in $K$ is within distance $\varepsilon$ of some point of $N$, i.e.
\[ \forall x \in K, \exists y \in N, \|x - y\| \leq \varepsilon. \]

LEMMA 2.8 ([22, Corollary 4.2.13]). There exists an $\varepsilon$-net with size \( \left( \frac{2}{\varepsilon} + 1 \right)^n \) of the unit $n$-Euclidean ball.

$\varepsilon$-nets can help us estimate the operator norm of a matrix.

LEMMA 2.9 ([22, Lemma 4.4.1, Exercise 4.4.3]). Let $A$ be an $m \times n$ matrix and $\varepsilon \in [0, 1)$. Then, for any $\varepsilon$-net $N$ of the unit sphere $S^{n-1}$, we have
\[ \sup_{x \in N} \|Ax\| \leq \|A\| \leq \frac{1}{1 - \varepsilon} \cdot \sup_{x \in N} \|Ax\|. \]

Moreover, if $m = n$ and $A$ is symmetric, we have
\[ \sup_{x \in N} |x^T Ax| \leq \|A\| \leq \frac{1}{1 - 2\varepsilon} \cdot \sup_{x \in N} |x^T Ax|. \]
2.3. Properties of random projections. Now we recall basic properties of random projection matrices. In this paper we call a matrix $P \in \mathbb{R}^{d \times n}$ a random projection matrix or a random matrix when its entries $P_{ij}$ are independently sampled from $N(0, 1/d)$.

One of the most important features of a random projection defined by a random matrix is that it nearly preserves the norm of any given vector with arbitrary high probability. The following lemma is known as a variant of the Johnson-Lindenstrauss lemma ([9]).

**Lemma 2.10 ( [22, Lemma 5.3.2, Exercise 5.3.3]).** Let $P \in \mathbb{R}^{d \times n}$ be a random matrix whose entries $P_{ij}$ are independently drawn from $N(0, 1/d)$. Then for any $x \in \mathbb{R}^n$ and $\varepsilon \in (0, 1)$, we have

$$\text{Prob} \left[ (1 - \varepsilon) \|x\|^2 \leq \|Px\|^2 \leq (1 + \varepsilon) \|x\|^2 \right] \geq 1 - 2 \exp(-C_0 \varepsilon^2 d),$$

where $C_0$ is an absolute constant.

Random projections also approximately preserve inner products, linear function values and quadratic function values.

**Lemma 2.11 ( [5, Lemma 3.1, 3.2, 3.3]).** Let $P \in \mathbb{R}^{d \times n}$ be a random matrix whose entries $P_{ij}$ are independently sampled from $N(0, 1/d)$. Then for any $x, y \in \mathbb{R}^n$, $A \in \mathbb{R}^{m \times n}$ having unit row vectors, $Q \in \mathbb{R}^{n \times n}$ and $\varepsilon \in (0, 1)$, the following probabilistic inequalities hold.

(i) With probability at least $1 - 4 \exp(-C_0 \varepsilon^2 d)$, we have

$$x^T y - \varepsilon \|x\| \|y\| \leq x^T P^T P y \leq x^T y + \varepsilon \|x\| \|y\| .$$

(ii) With probability at least $1 - 4m \exp(-C_0 \varepsilon^2 d)$,

$$Ax - \varepsilon \|x\| 1 \leq A P^T P x \leq Ax + \varepsilon \|x\| 1 .$$

(iii) With probability at least $1 - 8 \text{rank} Q \cdot \exp(-C_0 \varepsilon^2 d)$,

$$x^T Q x - 3\varepsilon \|x\|^2 \|Q\|_F \leq x^T P^T P Q P^T P x \leq x^T Q x + 3\varepsilon \|x\|^2 \|Q\|_F .$$

The above lemma, which estimates the error induced by random projections on different values, will be used to bound the error on the optimal value of a randomly projected quadratic optimization problem.

3. Convexified randomly projected problem.

3.1. Convexifying the objective function. In this section we give an error bound between $P$ and CRP. First we consider the distribution of the eigenvalues of $\bar{Q} = P Q P^T$. We can easily confirm that $\mathbb{E}[P Q P^T] = \frac{\mu Q}{d} I_d$

$$\mathbb{E}[P Q P^T]_{ij} = \mathbb{E} \left[ \sum_k \sum_l P_{ik} Q_{kl} P_{jl} \right] = \sum_k \sum_l Q_{kl} \mathbb{E} [P_{ik} P_{jl}] = \sum_k \sum_l Q_{kl} \frac{\delta_{ij} \delta_{kl}}{d} = \frac{\text{tr} Q}{d} \delta_{ij},$$
where $\delta_{ij}$ denotes the Kronecker delta symbol. By the above equality, we expect the eigenvalues of $PQP^T$ to be distributed around $\frac{\text{tr}Q}{d}$. One example of eigenvalue distributions of $Q$ and $PQP^T$ is shown in Figure 1, where we observe that the eigenvalue distribution of $PQP^T$ is skewed towards positive values and the negative spectrum of $PQP^T$ is negligible. In the next lemma, we evaluate the maximum deviation between the eigenvalues of $PQP^T$ and $\frac{\text{tr}Q}{d}$.

**Figure 1.** Distributions of eigenvalues of $Q$ and $PQP^T$ ($Q \in \mathbb{R}^{2000 \times 2000}$ with eigenvalues $\lambda_i(Q) = -10 + \frac{30i}{2000}, P \in \mathbb{R}^{200 \times 2000}$).

**Lemma 3.1.** Let $Q$ be an $n \times n$ symmetric matrix and let $P$ be a $d \times n$ random matrix whose entries are sampled from $N(0, 1/d)$. Then, for every $t \geq 0$,

$$
\text{Prob} \left[ \left\| PQP^T - \frac{\text{tr}Q}{d} I_d \right\| \geq t \right] \leq 2 \cdot 9^n \exp \left( -C_1 \min \left( \frac{d^2 t^2}{4C_6 \|Q\|_F^2}, \frac{d t}{2C_3 \|Q\|} \right) \right).
$$

**Proof.** First, using the eigenvalue decomposition of $Q$, we write $Q = U \Lambda U^T$, where $U$ is an orthogonal matrix and $\Lambda = \text{diag}(\lambda_1, \ldots, \lambda_n)$. Since the distribution of $PU$ is the same as the distribution of $P$, we have that the distribution of $PQP^T$ is the same as the distribution of $P \Lambda P^T$, hence

$$
\text{Prob} \left[ \left\| PQP^T - \frac{\text{tr}Q}{d} I_d \right\| \geq t \right] = \text{Prob} \left[ \left\| P\Lambda P^T - \frac{\text{tr}Q}{d} I_d \right\| \geq t \right].
$$

By Lemma 2.8, we can take a $1/4$-net $N$ with a size of $9^n$ of the unit $n$-Euclidean ball. Using Lemma 2.9 with $\varepsilon = 1/4$, we get

\begin{equation}
\text{Prob} \left[ \left\| P\Lambda P^T - \frac{\text{tr}Q}{d} I_d \right\| \geq t \right] \\
\leq \text{Prob} \left[ 2 \sup_{x \in N} \left| x^T (P\Lambda P^T - \frac{\text{tr}Q}{d} I_d) x \right| \geq t \right] \\
= \text{Prob} \left[ \sup_{x \in N} \left| x^T P\Lambda P^T x - \frac{\text{tr}Q}{d} \right| \geq \frac{t}{2} \right] \\
= \text{Prob} \left[ \sup_{x \in N} \left( \sum_{j=1}^n \left( \lambda_j(P^j, x)^2 - \frac{\lambda_j}{d} \right) \right) \geq \frac{t}{2} \right],
\end{equation}

where the $P^j$ are the column vectors of $P$.

Let $X_j = \langle P^j, x \rangle$, then $X_j$ are independent Gaussian random variables of variances $1/d$. Thus, by Lemma 2.2, we obtain

$$\|X_j\|_{\psi_2} \leq C \sqrt{\frac{t}{d}}.$$  

With the above inequality, Lemma 2.4 and Lemma 2.5, the random variable $\lambda_j X_j^2 - \frac{\lambda_j}{d}$ turns out to be sub-exponential whose sub-exponential norm is bounded by

$$\left\| \lambda_j X_j^2 - \frac{\lambda_j}{d} \right\|_{\psi_1} \leq C \left\| \lambda_j \|X_j\| \right\|_{\psi_1} \leq C \|\lambda_j\| X_j^2 \|_{\psi_2} \leq C \frac{\lambda_j}{d}.$$  

By Bernstein’s inequality (Theorem 2.6), we obtain, for each $x \in \mathcal{N}$, the following inequality:

$$(3.2) \quad \text{Prob} \left( \sum_{j=1}^n \left( \lambda_j X_j^2 - \frac{\lambda_j}{d} \right) \geq \frac{t}{2} \right) \leq 2 \exp \left( -C_1 \min \left( \frac{t^2}{4} \left( \frac{\lambda}{d} \right)^2, \frac{t}{2} \max \left( \frac{\lambda}{d} \right) \right) \right).$$

Finally, from (3.1), (3.2) and a union bound on $\mathcal{N}$, we have

$$\text{Prob} \left( \|PQPT - \frac{\text{tr} Q}{d} I_d\| \geq t \right) \leq 2 \cdot 9^d \exp \left( -C_1 \min \left( \frac{d^2 t^2}{4C^6 \|Q\|_F^2}, \frac{dt}{2C^3 \|Q\|} \right) \right),$$

which completes the proof.

**Corollary 3.2.** Let $Q$ be an $n \times n$ symmetric matrix and let $P$ be a $d \times n$ random matrix whose entries are sampled from $\text{N}(0, 1/d)$. If $\text{tr} Q > 0$, then for any $\varepsilon > 0$, we have

$$\text{Prob} \left( \min(0, \lambda_{\min}(PQPT)) \geq \varepsilon \|Q\|_F \right) \leq 2 \cdot 9^d \exp \left( -C_1 \min \left( \frac{(\text{tr} Q + \varepsilon d \|Q\|_F)^2}{4C^6 \|Q\|_F^2}, \frac{\text{tr} Q + \varepsilon d \|Q\|_F}{2C^3 \|Q\|} \right) \right),$$

where $\lambda_{\min}(M)$ denotes the minimum eigenvalue of a matrix $M$.

**Proof.** Suppose that $\|\min(0, \lambda_{\min}(PQPT))\| \geq \varepsilon \|Q\|_F$. This implies that $\lambda_{\min}(PQPT) < 0$ and $\lambda_{\min}(PQPT) \geq \varepsilon \|Q\|_F$. Furthermore, since $\text{tr} Q$ is positive, we have that $\lambda_{\min}(PQPT - \frac{\text{tr} Q}{d} I_d) < 0$ and $\lambda_{\min}(PQPT - \frac{\text{tr} Q}{d} I_d) \geq \frac{\text{tr} Q}{d} + \varepsilon \|Q\|_F$.

The last inequality implies that $\left\| PQPT - \frac{\text{tr} Q}{d} I_d \right\| \geq \frac{\text{tr} Q}{d} + \varepsilon \|Q\|_F$.

By the above argument, we obtain the following inequality:

$$\text{Prob} \left( \min(0, \lambda_{\min}(PQPT)) \geq \varepsilon \|Q\|_F \right) \leq \text{Prob} \left( \left\| PQPT - \frac{\text{tr} Q}{d} I_d \right\| \geq \frac{\text{tr} Q}{d} + \varepsilon \|Q\|_F \right).$$
Taking \( t = \frac{\text{tr} Q}{d} + \varepsilon \|Q\|_F \) in Lemma 3.1 ends the proof. \( \square \)

Next, we evaluate the difference between \( x^T Q x \) and \( x^T P^T \bar{Q}^+ P x \)
\((= x^T P^T \bar{F}^+ (PQP^T) P x)\) for a fixed vector \( x \), where \( \bar{F}^+ \) denotes the projection onto the positive semidefinite cone. The following theorem will be used to evaluate the error between the optimal values of \( \bar{P} \) and \( \bar{Q} \). In [5], the authors use Lemma 2.11 (iii) to evaluate the error between the optimal values of \( \bar{P} \) and \( \bar{Q} \). In this sense, Theorem 3.3 is an extension of Lemma 2.11 (iii).

The following theorem is proven by probabilistically evaluating the difference between the optimal values of \( \bar{R} \bar{P} \) and \( \bar{C} \bar{P} \) due to the randomness of problem \( \bar{R} \bar{P} \), though the convexification technique itself is a deterministic operation.

**Theorem 3.3.** Let \( Q \) be an \( n \times n \) symmetric matrix that satisfies \( \text{tr} \: Q > 0 \) and let \( P \) be a \( d \times n \) random matrix whose entries are sampled from \( N(0, 1/d) \).

Then, for any \( x \in \mathbb{R}^n \), with probability at least

\[
1 - 8 \text{rank} Q \cdot \exp(-C_0 \varepsilon_1^2 d) - 2 \exp(-C_0 \varepsilon_2^2 d) - 2 \cdot 9^d \exp \left( -C_1 \min \left( \frac{\left( \text{tr} Q + \varepsilon_3 d \|Q\|_F \right)^2}{4C^0 \|Q\|_F^2}, \frac{\text{tr} Q + \varepsilon_3 d \|Q\|_F}{2C^0 \|Q\|} \right) \right),
\]

we have

\[
| x^T Q x - x^T P^T \bar{F}^+ (PQP^T) P x | \leq (3 \varepsilon_1 + \varepsilon_3 + \varepsilon_2 \varepsilon_3) \|x\|^2 \|Q\|_F.
\]

**Proof.** By Lemma 2.10, Lemma 2.11 (iii) and Corollary 3.2, with probability at least

\[
1 - 8 \text{rank} Q \cdot \exp(-C_0 \varepsilon_1^2 d) - 2 \exp(-C_0 \varepsilon_2^2 d) - 2 \cdot 9^d \exp \left( -C_1 \min \left( \frac{\left( \text{tr} Q + \varepsilon_3 d \|Q\|_F \right)^2}{4C^0 \|Q\|_F^2}, \frac{\text{tr} Q + \varepsilon_3 d \|Q\|_F}{2C^0 \|Q\|} \right) \right),
\]

we have

\[
(3.3) \quad | x^T Q x - x^T P^T P Q P^T P x | \leq 3 \varepsilon_1 \|x\|^2 \|Q\|_F,
\]

\[
(3.4) \quad \|P x\|^2 \leq (1 + \varepsilon_2) \|x\|^2,
\]

\[
(3.5) \quad | \min(0, \lambda_{\text{min}}(PQP^T)) | \leq \varepsilon_3 \|Q\|_F.
\]

First, we decompose the error \( | x^T Q x - x^T P^T \bar{F}^+ (PQP^T) P x | \) into two terms:

\[
| x^T Q x - x^T P^T \bar{F}^+ (PQP^T) P x |
\leq | x^T Q x - x^T P^T P Q P^T P x |
+ | x^T P^T P Q P^T P x - x^T P^T \bar{F}^+ (PQP^T) P x |.
\]

The upper bound on the first term is given by (3.3). To bound the second term, we define \( \bar{Q}^- = PQP^T - \bar{F}^+ (PQP^T) \). Since \( -\bar{Q}^- \succeq 0 \), we can define its non-negative square root \( \sqrt{-\bar{Q}^-} \). With these notations, we get the upper bound on the second
term as follows:

\[
|x^T P^T P Q P^T P x - x^T P^T \mathcal{F}^+ (P Q P^T) P x| = |x^T P^T (-\tilde{Q}^-) P x|
\]

\[
= \left| \sqrt{-\tilde{Q}^-} P x \right|^2
\]

\[
\leq \left| \sqrt{-\tilde{Q}^-} \right|^2 \|P x\|^2
\]

\[
= \|\tilde{Q}^-\|\|P x\|^2
\]

\[
= \min(0, \lambda_{\min}(P Q P^T)) \|P x\|^2
\]

\[
\leq \varepsilon_3 \|Q\|_F (1 + \varepsilon_2)\|x\|^2.
\]

(by (3.4)(3.5))

In the next lemma we evaluate the probability shown in Corollary 3.2 or Theorem 3.3.

**Lemma 3.4.** Let \(Q\) be an \(n \times n\) symmetric matrix that satisfies \(\text{tr} \ Q > 0\) and define

\[
\tilde{r} \equiv \frac{\|Q\|_F^2}{\|Q\|^2}, \quad \tilde{k} \equiv \frac{\text{tr} \ Q}{\|Q\|}.
\]

Furthermore let \(D \geq C\).

If (i) \(d \geq -\frac{2 \tilde{k}}{\varepsilon_3 \sqrt{\tilde{r}}} + \frac{12D^6}{C_1 \varepsilon_3^2} - \frac{\log \delta}{3}\), and (ii) \(d < \frac{2D^3 \tilde{r} - \tilde{k}}{\varepsilon_3 \sqrt{\tilde{r}}}\), then

\[
2 \cdot 9^d \exp \left( -C_1 \min \left( \frac{(\text{tr} \ Q + \varepsilon_3 d \|Q\|_F)^2}{4C^6 \|Q\|^2_F}, \frac{\text{tr} \ Q + \varepsilon_3 d \|Q\|_F}{2C^3 \|Q\|} \right) \right) \leq \delta.
\]

**Proof.** For simplicity, we use \(p = \frac{\tilde{k}}{\sqrt{\tilde{r}}}\) and \(q = \sqrt{\tilde{r}}\). The conditions (i) and (ii) are equivalent to

\[
d \geq -\frac{2p}{\varepsilon_3} + \frac{12D^6}{C_1 \varepsilon_3^2} - \frac{\log \delta}{3},
\]

\[
d < \frac{2D^3 q - p}{\varepsilon_3}.
\]

We will prove that

\[
2 \cdot 9^d \exp \left( -C_1 \min \left( \frac{(\text{tr} \ Q + \varepsilon_3 d \|Q\|_F)^2}{4D^6 \|Q\|^2_F}, \frac{\text{tr} \ Q + \varepsilon_3 d \|Q\|_F}{2D^3 \|Q\|} \right) \right) \leq \delta
\]

holds which will end the proof as \(D \geq C\).
We can easily show that $2 \cdot 9^d \leq \exp(3d)$ for all $d \in \mathbb{N}$ and thus, we obtain

$$2 \cdot 9^d \exp \left( -C_1 \min \left( \frac{(\text{tr} \ Q + \varepsilon_3 d \|Q\|_F)^2}{4D^6 \|Q\|_F^2}, \frac{\text{tr} \ Q + \varepsilon_3 d \|Q\|_F}{2D^3 \|Q\|} \right) \right) \leq \exp \left( 3d - C_1 \min \left( \frac{(\text{tr} \ Q + \varepsilon_3 d \|Q\|_F)^2}{4D^6 \|Q\|_F^2}, \frac{\text{tr} \ Q + \varepsilon_3 d \|Q\|_F}{2D^3 \|Q\|} \right) \right)$$

which can be easily verified by squaring both sides. Applying this inequality

$$\leq \exp \left( 3d - \frac{C_1}{4D^6} (p + \varepsilon_3 d)^2 \right).$$

Since (3.7) is equivalent to $p + \varepsilon_3 d < 2D^3 q$, we have

$$\min \left( \frac{(p + \varepsilon_3 d)^2}{4D^6}, \frac{pq + \varepsilon_3 dq}{2D^3} \right) = \frac{(p + \varepsilon_3 d)^2}{4D^6}.$$  

Thus,

$$2 \cdot 9^d \exp \left( -C_1 \min \left( \frac{(\text{tr} \ Q + \varepsilon_3 d \|Q\|_F)^2}{4D^6 \|Q\|_F^2}, \frac{\text{tr} \ Q + \varepsilon_3 d \|Q\|_F}{2D^3 \|Q\|} \right) \right) \leq \exp \left( 3d - \frac{C_1}{4D^6} (p + \varepsilon_3 d)^2 \right).$$

Next, we show that $3d - \frac{C_1}{4D^6} (p + \varepsilon_3 d)^2 = \log \delta$, which will complete the proof. Note that the quadratic equation,

$$3d - \frac{C_1}{4D^6} (p + \varepsilon_3 d)^2 = \log \delta$$

is equivalent to

$$C_1 \varepsilon_3^2 d^2 + (2C_1 \varepsilon_3 p - 12D^6) d + 4D^6 \log \delta + C_1 p^2 = 0$$

of which real solutions are given by (if there are any)

$$d = \frac{-2C_1 \varepsilon_3 p + 12D^6 \pm \sqrt{(2C_1 \varepsilon_3 p - 12D^6)^2 - 4C_1 \varepsilon_3^2 (4D^6 \log \delta + C_1 p^2)}}{2C_1 \varepsilon_3}$$

$$= \frac{-C_1 \varepsilon_3 p + 6D^6 \pm \sqrt{36D^{12} - 12D^6 C_1 \varepsilon_3 p - 4D^6 C_1 \varepsilon_3^2 \log \delta}}{C_1 \varepsilon_3^3}.$$  

If there are no real roots, then $3d - \frac{C_1}{4D^6} (p + \varepsilon_3 d)^2 < \log \delta$ holds for all $d$. Thus, it is sufficient to show that

$$d \geq \frac{-C_1 \varepsilon_3 p + 6D^6 \pm \sqrt{36D^{12} - 12D^6 C_1 \varepsilon_3 p - 4D^6 C_1 \varepsilon_3^2 \log \delta}}{C_1 \varepsilon_3^3}.$$  

To show this inequality, we use the following inequality $a - \frac{b}{2a} \geq \sqrt{a^2 - b^2}$ $(a, b > 0, a^2 > b)$ which can be easily verified by squaring both sides. Applying this inequality with $a = 6D^6$ and $b = 12D^6 C_1 \varepsilon_3 p + 4D^6 C_1 \varepsilon_3^2 \log \delta$, we obtain

$$\text{(3.8) } 6D^6 - \frac{12D^6 C_1 \varepsilon_3 p + 4D^6 C_1 \varepsilon_3^2 \log \delta}{2 \cdot 6D^6} \geq \sqrt{36D^{12} - 12D^6 C_1 \varepsilon_3 p - 4D^6 C_1 \varepsilon_3^2 \log \delta}.$$
which completes the proof:

\[
\begin{align*}
\epsilon \geq -\frac{2p}{\epsilon_3^2} + \frac{12D^6}{C_1\epsilon_3^2} - \frac{\log \delta}{3} & \quad \text{(by (3.6))} \\
= \frac{-C_1\epsilon_3^2 + 6D^6}{\epsilon_3^2} + \frac{1}{C_1\epsilon_3^2} \left( 6D^6 - \frac{12D^6C_1\epsilon_3^2 + 4D^6C_1\epsilon_3^2 \log \delta}{2 \cdot 6D^6} \right) & \\
\geq \frac{-C_1\epsilon_3^2 + 6D^6 + \sqrt{36D^{12} - 12D^{12}C_1\epsilon_3^2 - 4D^{12}C_1\epsilon_3^2 \log \delta}}{C_1\epsilon_3^2} & \quad \text{(by (3.8))}
\end{align*}
\]

Remark 3.5. The quantities \( \hat{\epsilon} \) and \( \hat{k} \) defined in Lemma 3.4 are known as stable rank (also called numerical rank) [10] and effective rank (also called intrinsic dimension) [20, 21]. Clearly, \( \hat{\epsilon}, \hat{k} \leq \text{rank} \, Q \) and they can be interpreted as the robust version of the usual rank. These quantities are used in covariance estimation [22].

3.2. The error bound in a special case. We now evaluate the error between the optimal value of the original problem (1.1) and that of the convexified projected problem (1.3) under the following assumptions.

Assumption 3.6. In the original problem (1.1), we assume the followings:

(A1) The problem (1.1) has a finite optimal value at \( x = x^\ast \).

(A2) All the rows of \( A \) are unit vectors (i.e., \( \|A_i\| = 1 \)).

(A3) \( \text{tr} \, Q > 0 \).

(A4) There exists a closed ball \( B(0, r) \) which is contained in the polytope \( Ax \leq b \).

(A2) holds without loss of generality: if \( \|A_i\| \neq 1 \), we replace \( A_i \) by \( A_i/\|A_i\| \) and \( b_i \) by \( b_i/\|A_i\| \) and then the assumption is satisfied. (A3) is essential in this paper, though it is replaced by a weaker assumption later in section 4. As shown in Lemma 3.1 and Corollary 3.2, the eigenvalues of \( Q = PPP^T \) concentrate around \( \frac{\text{tr} \, Q}{\text{dim}} \), and ignoring the negative spectrum of \( PPP^T \) does not change the problem so much especially when \( \text{tr} \, Q \) is large enough. (A4) also can be weaken later in subsection 3.3; An essential requirement is that the polytope is full dimensional; equality constraints are not acceptable in (1.1).

We investigate the relationship between \( P, \, CRP, \) and the following problem:

\[
\text{CRP}_\varepsilon \equiv \min \{ u^T \hat{Q}^+ u + \hat{c}^T u \mid \hat{A} u \leq b + \varepsilon \|x^\ast\| \, 1 \},
\]

where \( \varepsilon > 0 \) and \( 1 \) is the all-one vector.

For an optimization problem \( F \), we denote by \( \text{opt}(F) \) the optimal objective function value of \( F \). We also let \( S, T \) and \( T_\varepsilon \) be the feasible regions of \( P, \, CRP \) and \( \text{CRP}_\varepsilon \), respectively.

We can easily show that \( \text{opt}(\text{CRP}) \geq \text{opt}(P) \) and \( \text{opt}(\text{CRP}) \geq \text{opt}(\text{CRP}_\varepsilon) \).

Lemma 3.7. Under Assumption 3.6 (A1) and (A4), \( \text{opt}(\text{CRP}) \) is finite and

\[
\text{opt}(\text{CRP}) \geq \text{opt}(P).
\]

Proof. First, we show that \( \text{opt}(\text{CRP}) \) is finite. Assumption 3.6 (A4) implies that \( u = 0 \) is feasible for \( \text{CRP} \), and therefore, \( \text{opt}(\text{CRP}) < \infty \). We can confirm that \( \text{opt}(\text{CRP}) > -\infty \) by contradiction. If \( \text{opt}(\text{CRP}) = -\infty \), there exists a sequence \( \{u^k\}_{k=1}^\infty \) in \( T \) such that \( u^k^T \hat{Q}^+ u^k + \hat{c}^T u^k \to -\infty \) \( (k \to \infty) \). Since \( u^k \in T \), we have
\[ \bar{A}u^k = AP^Tu^k \leq b, \text{ which implies } P^Tu^k \in S \text{ and} \]
\[
\text{opt}(P) \leq (P^Tu^k)^TQ(P^Tu^k) + c^T(P^Tu^k) \\
\leq u^kT \mathcal{F}^+(PQP^T)u^k + (Pc)^Tu^k \\
= u^kT \bar{Q}^+ u^k + c^Tu^k \to -\infty,
\]
which implies a contradiction to Assumption 3.6 (A1).

To show the rest part of the proof, let \( u^* \) be an optimum of CRP. From the same argument before, we have that \( P^Tu^* \in S \) and
\[
\text{opt}(P) \leq (P^Tu^*)^TQ(P^Tu^*) + c^T(P^Tu^*) \\
\leq u^*T \bar{Q}^+ u^* + c^Tu^* \\
= \text{opt(CRP)}. \]

**Lemma 3.8.** Under Assumption 3.6 (A1) and (A4),
\[
\text{opt(CRP)} \geq \text{opt(CRP}_\varepsilon).
\]

**Proof.** This follows immediately from \( T \subseteq T_\varepsilon \) of two feasible regions. Note that, as shown in Lemma 3.7, we have \( 0 \in T \) and the feasibilities of CRP and CRP\(_\varepsilon\) are guaranteed.

The inequality \( \text{opt(CRP)} \geq \text{opt(CRP}_\varepsilon \) includes the case where \( \text{opt(CRP}_\varepsilon \) = \( -\infty \), though, as we will see later in Theorem 3.10, this case does not occur.

Next, we investigate the gap between \( P \) and CRP\(_\varepsilon\).

**Theorem 3.9.** Let \( x^* \in \mathbb{R}^n \) be an optimum of \( P \). Under Assumption 3.6 (A1), (A2) and (A3), for any \( \varepsilon, \varepsilon_1, \varepsilon_2, \varepsilon_3, \varepsilon_4 > 0 \), we have
\[
\text{opt}(P) \geq \text{opt(CRP}_\varepsilon) - (3\varepsilon_1 + \varepsilon_3 + \varepsilon_2\varepsilon_3) \|x^*\|^2 \|Q\|_F - \varepsilon_4 \|x^*\| \|c\|
\]
with probability at least
\[
1 - 4m \exp(-C_0\varepsilon^2d) - 8 \text{ rank } Q \cdot \exp(-C_0\varepsilon_1^2d) - 2 \exp(-C_0\varepsilon_2^2d) - 2 \exp(-C_0\varepsilon_4^2d) \\
- 2 \cdot 9^d \exp \left(-C_1 \min \left( \frac{\text{tr } Q + \varepsilon_3d \|Q\|_F^2}{4C^6 \|Q\|_F^2}, \frac{\text{tr } Q + \varepsilon_3d \|Q\|_F}{2C^3 \|Q\|} \right) \right).
\]

**Proof.** By Lemma 2.11(i)(ii) and Theorem 3.3, with probability at least
\[
1 - 4m \exp(-C_0\varepsilon^2d) - 8 \text{ rank } Q \cdot \exp(-C_0\varepsilon_1^2d) - 2 \exp(-C_0\varepsilon_2^2d) - 2 \exp(-C_0\varepsilon_4^2d) \\
- 2 \cdot 9^d \exp \left(-C_1 \min \left( \frac{\text{tr } Q + \varepsilon_3d \|Q\|_F^2}{4C^6 \|Q\|_F^2}, \frac{\text{tr } Q + \varepsilon_3d \|Q\|_F}{2C^3 \|Q\|} \right) \right),
\]
the following inequalities hold:
\[
\begin{align*}
\text{(3.9)} & \quad \quad AP^Tx^* \leq Ax^* + \varepsilon \|x^*\|, \\
\text{(3.10)} & \quad \quad |x^TQx^* - x^TP^TF^+(PQP^T)P^Tx^*| \leq (3\varepsilon_1 + \varepsilon_3 + \varepsilon_2\varepsilon_3) \|x^*\|^2 \|Q\|_F, \\
\text{(3.11)} & \quad \quad c^TP^Tx^* \leq c^Tx^* + \varepsilon_4 \|x^*\| \|c\|.
\end{align*}
\]
(3.9) implies that $Px^*$ is a feasible solution for $\text{CRP}_\varepsilon$. By (3.10) and (3.11), we have

$$
\text{opt}(P) = x^TQx^* + c^Tx^*
$$

$$
\geq x^T P^T F^+ (PQP^T) Px^* - (3\varepsilon_1 + \varepsilon_3 + \varepsilon_2\varepsilon_3) \|x^*\|^2 \|Q\|_F + c^T P^T Px^* - \varepsilon_4 \|x^*\| \|c\|\n$$

$$
= (Px^*)^T Q^+(Px^*) + c^T(Px^*) - (3\varepsilon_1 + \varepsilon_3 + \varepsilon_2\varepsilon_3) \|x^*\|^2 \|Q\|_F - \varepsilon_4 \|x^*\| \|c\|.
$$

$$
\geq \text{opt}(\text{CRP}_\varepsilon) - (3\varepsilon_1 + \varepsilon_3 + \varepsilon_2\varepsilon_3) \|x^*\|^2 \|Q\|_F - \varepsilon_4 \|x^*\| \|c\|. \qed
$$

It should be noted that the error estimate in Theorem 3.9 includes the information on the optimum of $P$ (more precisely, $\|x^*\|$). The estimate seems unrealistic because $x^*$ is not available. However, the necessary quantity is $\|x^*\|$. In the case of a bounded feasible region for $P$, it may be bounded by some threshold $R$ such as $\|x^*\| \leq R$ by proving that the feasible region of $P$ is included in a ball $\|x\| \leq R$.

**Theorem 3.10.** Under Assumption 3.6 (A1), (A2) and (A4), $\text{opt}(\text{CRP}_\varepsilon)$ is finite and we have

$$
\text{opt}(\text{CRP}) \geq \text{opt}(\text{CRP}_\varepsilon) \geq 1 + \varepsilon \frac{\|x^*\|}{r} \text{opt}(\text{CRP}).
$$

*Proof.* First, we admit the finiteness of $\text{opt}(\text{CRP}_\varepsilon)$ and show the inequality. We will confirm the finiteness at the end of this proof.

Since we have already shown that $\text{opt}(\text{CRP}) \geq \text{opt}(\text{CRP}_\varepsilon)$ in Lemma 3.8, the second inequality:

$$
\text{opt}(\text{CRP}_\varepsilon) \geq \frac{\text{opt}(\text{CRP})}{\alpha},
$$

where we let $\alpha = 1 - \frac{\varepsilon \|x^*\|}{r + \varepsilon \|x^*\|}$, is proven now. Since $B(0, r) \subset S$, $Ax \leq b$ for all $x$ satisfying $\|x\| \leq r$. Therefore,

$$
(3.12) \quad b_i \geq \max_{\|x\| \leq r} (Ax)_i = r \|A_i\| = r \quad (i = 1, 2, \ldots, m).
$$

Next, we construct a feasible solution for $\text{CRP}$ close to $u^*_\varepsilon$, an optimum of $\text{CRP}_\varepsilon$. This will enable us to evaluate the error between $\text{opt}(\text{CRP})$ and $\text{opt}(\text{CRP}_\varepsilon)$.

By the definition of $\text{CRP}_\varepsilon$, we have $\bar{A}u^*_\varepsilon \leq b + \varepsilon \|x^*\| \mathbf{1}$, and thus, for each $i = 1, 2, \ldots, m$, we have

$$
(\alpha\bar{A}u^*_\varepsilon)_i \leq \left(1 - \frac{\varepsilon \|x^*\|}{r + \varepsilon \|x^*\|}\right) (b_i + \varepsilon \|x^*\|)
$$

$$
= b_i + \varepsilon \|x^*\| - \frac{\varepsilon \|x^*\|}{r + \varepsilon \|x^*\|} (b_i + \varepsilon \|x^*\|)
$$

$$
\leq b_i + \varepsilon \|x^*\| - \frac{\varepsilon \|x^*\|}{r + \varepsilon \|x^*\|} (r + \varepsilon \|x^*\|) \quad \text{(by (3.12))}
$$

$$
= b_i.
$$

This implies that $\alpha u^*_\varepsilon$ is feasible for $\text{CRP}$, hence

$$
(3.13) \quad \text{opt}(\text{CRP}) \leq g(\alpha u^*_\varepsilon),
$$

where $g(x) = x^T P^T F^+ F P x + c^T P x - \varepsilon_4 \|x\| \|c\|$.
where \(g(\cdot)\) is the objective function of CRP and CRP\(_\varepsilon\), i.e. \(g(u) = u^TQ^+u + \tilde{c}^Tu\). Note that \(Q^+\) is positive semidefinite, which implies \(g\) is convex and thus we have

\[
g(au^*_\varepsilon) = g(au^*_\varepsilon + (1 - \alpha)0) \leq \alpha g(u^*_\varepsilon) + (1 - \alpha)g(0) = \alpha \text{opt(CRP)}.
\]

From the two inequalities (3.13) and (3.14),

\[
\text{opt(CRP)} \geq \frac{\alpha \text{opt(CRP)}}{\alpha}
\]

is derived.

And lastly, we show that opt(CRP\(_\varepsilon\)) is finite. Let \(u_\varepsilon\) be an arbitrary feasible solution of CRP\(_\varepsilon\), i.e. \(Au_\varepsilon \leq b + \varepsilon \|x^*\| \mathbf{1}\), then, substituting \(u_\varepsilon\) for \(u^*_\varepsilon\) in the above discussion, we have

\[
\text{opt(CRP)} \leq g(au_\varepsilon) \leq \alpha g(u_\varepsilon).
\]

Since we have already shown the finiteness of opt(CRP) in Lemma 3.7, \(g(u_\varepsilon)\) is turned out to be lower bounded, and thus opt(CRP\(_\varepsilon\)) is finite.

**Lemma 3.11.** Under Assumption 3.6, for any \(\varepsilon, \varepsilon_1, \varepsilon_2, \varepsilon_3, \varepsilon_4 > 0\), we have

\[
\text{opt(CRP)} \geq \text{opt(P)}
\]

\[
\geq \left(1 + \varepsilon \frac{\|x^*\|}{r}\right) \text{opt(CRP)} - (3\varepsilon_1 + \varepsilon_3 + \varepsilon_2\varepsilon_3) \|x^*\|^2 \|Q\|_F - \varepsilon_4 \|x^*\| \|c\|
\]

with probability at least

\[
1 - 4m \exp(-C_0\varepsilon^2d) - 8 \text{rank}Q \cdot \exp(-C_0\varepsilon^2d) - 2 \exp(-C_0\varepsilon_2^2d) - 2 \exp(-C_0\varepsilon_4^2d)
\]

\[
- 2 \cdot 9^d \exp\left(-C_1 \min\left\{\frac{(\text{tr}Q + \varepsilon_3d \|Q\|_F)^2}{4C^6 \|Q\|_F^2}, \frac{\text{tr}Q + \varepsilon_3d \|Q\|_F}{2C^3 \|Q\|}\right\}\right).
\]

**Proof.** This statement follows from Lemma 3.7, Theorem 3.9 and Theorem 3.10.

**Theorem 3.12** (Approximation theorem under Assumption 3.6). Define \(\hat{r} \equiv \frac{\|Q\|_F^2}{\|Q\|^2}\), \(\hat{k} \equiv \frac{\text{tr}Q}{\|Q\|}\). Let \(\varepsilon, \varepsilon_1, \varepsilon_3, \varepsilon_4 \in (0, 1], \delta_1, \delta_2 \in (0, 1/2)\) and \(D \geq C\). Suppose that

(i) \(d \geq \max\left\{\frac{\log((24 \text{rank}Q + 6)/\delta_1)}{C_0\varepsilon_1^2}, \frac{\log(6/\delta_1)}{C_0\varepsilon_2^2}, \frac{\log(12m/\delta_1)}{C_0\varepsilon_4^2}\right\}\)

(ii) \(d \geq -\frac{2}{\varepsilon_3 \sqrt{\hat{r}}} + \frac{12D^6}{C_1\varepsilon_3^2} - \frac{\log \delta_2}{3}\)

(iii) \(d < \frac{2D^3\hat{r} - \hat{k}}{\varepsilon_3 \sqrt{\hat{r}}}\).

Then with probability at least \(1 - \delta_1 - \delta_2\), the following inequalities hold:

\[
\text{opt(CRP)} \geq \text{opt(P)}
\]

\[
\geq \left(1 + \varepsilon \frac{\|x^*\|}{r}\right) \text{opt(CRP)} - (3\varepsilon_1 + 2\varepsilon_3) \|x^*\|^2 \|Q\|_F - \varepsilon_4 \|x^*\| \|c\|
\]

**Proof.** Let \(\varepsilon_2 = 1\) in Lemma 3.11, then we have

\[
\text{opt(CRP)} \geq \text{opt(P)}
\]

\[
\geq \left(1 + \varepsilon \frac{\|x^*\|}{r}\right) \text{opt(CRP)} - (3\varepsilon_1 + 2\varepsilon_3) \|x^*\|^2 \|Q\|_F - \varepsilon_4 \|x^*\| \|c\|
\]
with probability at least

\[ 1 - 4m \exp(-C_0 \varepsilon^2 d) - 8 \text{rank} Q \cdot \exp(-C_0 \varepsilon_1^2 d) - 2 \exp(-C_0 d) - 2 \exp(-C_0 \varepsilon^2 d) \\
- 2 \cdot 9^d \exp \left( -C_1 \min \left( \frac{\text{tr} Q + \varepsilon_3 d \|Q\|_F^2}{4C_0 \|Q\|_F^2}, \frac{\text{tr} Q + \varepsilon_3 d \|Q\|_F}{2C^3 \|Q\|} \right) \right). \]

By (i) and the assumption \( \varepsilon_1 \leq 1 \),

\[ 8 \text{rank} Q \cdot \exp(-C_0 \varepsilon_1^2 d) + 2 \exp(-C_0 d) \leq (8 \text{rank} Q + 2) \exp(-C_0 \varepsilon_1^2 d) \leq \frac{\delta_1}{3}. \]

(i) also implies

\[ 2 \exp(-C_0 \varepsilon^2 d) \leq \frac{\delta_1}{3}, \]

\[ 4m \exp(-C_0 \varepsilon^2 d) \leq \frac{\delta_1}{3}. \]

Since (ii), (iii) are the same assumptions as the ones made in Lemma 3.4, we have

\[ 2 \cdot 9^d \exp \left( -C_1 \min \left( \frac{\text{tr} Q + \varepsilon_3 d \|Q\|_F^2}{4C_0 \|Q\|_F^2}, \frac{\text{tr} Q + \varepsilon_3 d \|Q\|_F}{2C^3 \|Q\|} \right) \right) \leq \delta_2, \]

ending the proof.

Theorem 3.12 has shown some upper and lower bounds on the size \( d \) of a randomly projected QP. Indeed, to bound \(|\text{opt}(P) - \text{opt}(CRP)|\), the existence of these bounds is reasonable. Larger \( d \) makes the gap between \( \text{opt}(P) \) and \( \text{opt}(RP) \) small, while smaller \( d \) makes the gap between \( \text{opt}(RP) \) and \( \text{opt}(CRP) \) small because the matrix \( \bar{Q} \approx \text{tr} Q d I_d \) in the objective of \( \text{opt}(RP) \) tends to be positive definite. Therefore, to make the bound of \(|\text{opt}(P) - \text{opt}(CRP)|\) small, a well-balanced \( d \) for both gaps is needed. In Proposition 3.15, we show that for a certain class of non-convex QPs (1.1), there exists \( d \) satisfying the above (i)-(iii), and furthermore, how small \( d \) can be in those cases.

Now we discuss how to obtain a feasible solution for the original problem \( P \) from \( u^* \), the optimal solution of \( CRP \). As shown in the proof of Lemma 3.7, \( x' = P^T u^* \) is feasible for \( P \) and

\[ \text{opt}(CRP) \geq x'^T Q x' + c^T x' \geq \text{opt}(P), \]

so that \( x' \) is a feasible solution of \( P \) that we expect to achieve an approximate optimal value.

3.3. The error bound in a more general case. Previously, we assumed that the feasible region contains a sphere centered at the origin ((A4) in Assumption 3.6). Next, we will consider a more general situation, i.e. we make the following assumption.

Assumption 3.13. (A1),(A2),(A3) in Assumption 3.6 and
(A4') There exists a closed ball \( B(x_0, r) \) which is contained in the polytope \( Ax \leq b \).

Considering the variable translation \( y = x - x_0 \), we obtain the translated problem:

\[ P_T \equiv \min_y \{ y^T Q y + (2Qx_0 + c)^T y \mid A y \leq b - Ax_0 \}. \]
It is obvious that \( \text{opt}(P_T) = \text{opt}(P) - x_0^T Q x_0 - c^T x_0 \), hence it is enough to solve \( P_T \) instead of \( P \). Moreover, there exists a closed ball \( B(0, r) \) which is contained in the polytope \( Ax \leq b - Ax_0 \) so that we can apply the previous argument. Define the convexified randomly projected problem of \( P_T \):

\[
\text{CRP}_T \equiv \min_v \{ v^T Q^+ v + (P(2Qx_0 + c))^T v \mid \tilde{A}v \leq b - Ax_0 \},
\]

and then, by Theorem 3.12, we obtain a generalized approximation theorem. Note that one of optimal points of \( P_T \) is \( y^* = x^* - x_0 \).

**Theorem 3.14 (Approximation theorem under Assumption 3.13).** Define \( \tilde{r} \equiv \frac{\|Q\|_F^2}{\|Q\|^2}, \tilde{k} \equiv \frac{\text{tr} Q}{\|Q\|} \). Let \( \varepsilon, \varepsilon_1, \varepsilon_2 \in (0, 1], \delta_1, \delta_2 \in (0, 1/2) \) and \( D \geq C \).

Suppose that

1. \( d \geq \max \left\{ \frac{\log((24 \text{ rank } Q + 6)/\delta_1)}{C_0\varepsilon_1^2}, \frac{\log(6/\delta_1)}{C_0\varepsilon_2^2}, \frac{\log(12m/\delta_1)}{C_0\varepsilon_4^2} \right\} \),
2. \( d \geq -\frac{2\tilde{k}}{\varepsilon_3 \sqrt{\tilde{r}}} + \frac{12D^6}{C_1\varepsilon_3^2} - \frac{\log \delta_2}{3}, \)
3. \( d < \frac{2D^3 \tilde{r} - \tilde{k}}{\varepsilon_3 \sqrt{\tilde{r}}} \).

Then with probability at least \( 1 - \delta_1 - \delta_2 \), the following inequality holds:

\[
\text{opt(\text{CRP}_T)} \geq \text{opt}(P_T)
\geq \left( 1 + \varepsilon \frac{\|y^*\|}{r} \right) \text{opt(\text{CRP}_T)} - (3\varepsilon_1 + 2\varepsilon_2) \|y^*\|^2 \|Q\|_F
\geq \varepsilon_4 \|y^*\| \|2Qx_0 + c\|.
\]

Now we show some conditions on the size \( n, m \) and \( \text{cond}(Q) \) for non-convex QP (1.1) to guarantee the existence of \( d \) satisfying the above (i)-(iii) in Theorem 3.14, where we define \( \text{cond}(Q) \equiv \sigma_n/\sigma_1 \), the condition number of \( Q \). In the following proposition, \( P(n) \) may be formed with specific functions such as \( \log^3 n \) and \( n^7 (0 < \tau < 1) \).

**Proposition 3.15.** Assume that \( m < \tilde{C} n \) holds for some constant \( \tilde{C} \) and a function \( P(n) \) satisfies that

\[
\log n \ll P(n) \quad \text{and} \quad P(n) + \sqrt{n} \ll \frac{\sqrt{n} P(n)}{\text{cond}(Q)^2}.
\]

Then, if \( n \) is large enough, we can choose

\[
d = O \left( \frac{P(n)}{\varepsilon_0^2} \right)
\]

such that the (i)-(iii) of Theorem 3.14 are satisfied. Here, \( \varepsilon_0 \) is a constant that only depends on \( \varepsilon, \varepsilon_1, \varepsilon_4, \delta_1 \).

**Proof.** By (i) and the assumption \( \log n \ll P(n) \), we can take \( d = \frac{P(n)}{\varepsilon_0^2} \) using a constant \( \varepsilon_0 \) that only depends on \( \varepsilon, \varepsilon_1, \varepsilon_4, \delta_1 \).

We can choose \( D^6 = C' \left( d + \frac{\tilde{k}}{\sqrt{\tilde{r}}} \right) = C' \left( \frac{P(n)}{\varepsilon_0^2} + \frac{\tilde{k}}{\sqrt{\tilde{r}}} \right) \), where \( C' \) is a constant, that depends only on \( \varepsilon, \delta_2 \), such that (ii) is satisfied.
To finish the proof we need to verify (iii):

\[
P(n) - \hat{k} < \frac{2\sqrt{C'} \left( \frac{P(n)}{\varepsilon_0} + \frac{\hat{k}}{\varepsilon_0} \right)}{\varepsilon_3 \sqrt{r}},
\]

which is equivalent to:

\[
(3.15) \quad \frac{\varepsilon_3}{\varepsilon_0} \sqrt{n}P(n) + \hat{k} < 2\sqrt{C'} \sqrt{\frac{P(n)}{\varepsilon_0} \hat{r}^2 + \frac{\hat{k} \hat{r}^{3/2}}{2}}.
\]

From the definitions of \( \hat{r} \) and \( \hat{k} \), we easily see that

\[
(3.16) \quad \frac{1}{\text{cond}(Q)^2} n \leq \hat{r} \leq n \quad \text{and} \quad 0 < \hat{k} < n.
\]

Hence, the left-hand-side, \( \frac{\varepsilon_3}{\varepsilon_0} \sqrt{n}P(n) + \hat{k} \), of (3.15) has

\[
\frac{\varepsilon_3}{\varepsilon_0} \sqrt{n}P(n) + n
\]

as an upper bound, for \( n \) large enough, and the right-hand-side of (3.15) is lower-bounded by

\[
\frac{2\sqrt{C'}}{\varepsilon_0 \text{cond}(Q)^2} n \sqrt{P(n)}.
\]

Hence the condition

\[
P(n) + \sqrt{n} \ll \frac{\sqrt{nP(n)}}{\text{cond}(Q)^2}
\]

is enough to prove (iii) and hence that \( d = \frac{P(n)}{\varepsilon_0} \) satisfies the three condition of Theorem 3.14, for \( n \) large enough. \(\square\)

3.4. Relative error. The approximation inequality shown so far has an additive form. But we do not know how large or small the error term appearing in the theorem is compared to the optimal value \( \text{opt}(P_T) \). The purpose of this section is to transform the approximation inequality in Theorem 3.14 into a multiplicative form, \( \eta \cdot \text{opt}(P_T) \geq \text{opt}(CRP_T) \geq \text{opt}(P_T) \), where \( \eta \geq 1 \) measures how much the error term is relative to the optimal value. Writing the approximation as above allows to see the parameters of \( P_T \) that influence the relative error between the two problems.

For a matrix \( M \in \mathbb{R}^{n \times n} \) and a vector \( w \in \mathbb{R}^n \), we treat \( (M, w) \) as an \( (n^2 + n) \)-dimensional vector where the inner product of \( (M_1, w_1) \) and \( (M_2, w_2) \) is defined as follows:

\[
((M_1, w_1), (M_2, w_2))_{\mathbb{R}^{n^2+n}} \equiv \text{tr}(M_1^T M_2) + w_1^T w_2.
\]

**Corollary 3.16.** We can rewrite the approximation inequality shown in Theorem 3.14 into the multiplicative one:

\[
\left( \frac{r}{r + \varepsilon \|y^*\|} \right) \left( 1 + \frac{[(3\varepsilon_1 + 2\varepsilon_3)^2 + \varepsilon_4^2]}{\cos \theta^*} \right) \text{opt}(P_T) \geq \text{opt}(\text{CRP}_T) \geq \text{opt}(P_T)
\]
where \( \theta^* \) is the angle between vectors \( \xi \equiv (y^* y^\top, y^*) \) and \( \zeta \equiv (Q, 2Qx_0 + c) \in \mathbb{R}^{n^2+n} \).

**Proof.** By the definition of \( \xi \) and \( \zeta \), we have,

\[
\text{opt}(P_T) = \langle \xi, \zeta \rangle_{\mathbb{R}^{n^2+n}}.
\]

We also have

\[
\|\xi\|^2 \|\zeta\|^2 = (\|y^* y^\top\|_F^2 + \|y^*\|^2)(\|Q\|_F^2 + \|2Qx_0 + c\|^2)
\]

\[
= (\|y^*\|^4 + \|y^*\|^2)(\|Q\|_F^2 + \|2Qx_0 + c\|^2)
\]

\[
\geq \|y^*\|^4 \|Q\|_F^2 + \|y^*\|^2 \|2Qx_0 + c\|^2,
\]

and now we can evaluate the error term in Theorem 3.14:

\[
E \equiv (3\varepsilon_1 + 2\varepsilon_3) \|y^*\|^2 \|Q\|_F + \varepsilon_4 \|y^*\| \|2Qx_0 + c\|
\]

\[
\leq \sqrt{((3\varepsilon_1 + 2\varepsilon_3)^2 + \varepsilon_4^2)(\|y^*\|^4 \|Q\|_F^2 + \|y^*\|^2 \|2Qx_0 + c\|^2)}
\]

(by Schwarz’s inequality)

\[
= \sqrt{((3\varepsilon_1 + 2\varepsilon_3)^2 + \varepsilon_4^2) \|\xi\| \|\zeta\|}
\]

\[
= \sqrt{((3\varepsilon_1 + 2\varepsilon_3)^2 + \varepsilon_4^2) \frac{\langle \xi, \zeta \rangle_{\mathbb{R}^{n^2+n}}}{\cos \theta^*}}
\]

\[
= \sqrt{((3\varepsilon_1 + 2\varepsilon_3)^2 + \varepsilon_4^2) \text{opt}(P_T)} / \cos \theta^*.
\]

\[
4. \textbf{Scaling and Preconditioning.} \text{ In this section we will provide an error bound for CRP under a weaker assumption than (A3) in Assumption 3.6.}
\]

**Assumption 4.1.** (A1),(A2) in Assumption 3.6, (A4') in Assumption 3.13 and (A3') At least one of the eigenvalues of \( Q \) is positive.

We consider the scaling \( y = Uz \) and the following scaled problem:

\[
P'_T \equiv \min_{z} \{ z^\top Q' z + (U^\top (2Qx_0 + c))^\top z \mid A'z \leq b - Ax_0 \},
\]

where \( U \) is a scaling invertible matrix, \( Q' = U^\top QU \) and \( A' = AU \). Obviously, we have \( \text{opt}(P'_T) = \text{opt}(P_T) \). The corresponding convexified randomly projected problem becomes

\[
\text{CRP}'_T \equiv \min_{w} \{ w^\top Q'^\top w + (PU^\top (2Qx_0 + c))^\top w \mid A'w \leq b - Ax_0 \},
\]

and \( \text{opt}(\text{CRP}'_T) = \text{opt}(\text{CRP}_T) \) holds.

In order to apply the arguments so far, we have to make sure that \( Q' \) is positive. Von Neumann’s trace inequality [16] implies that

\[
\text{tr} Q' = \text{tr}(U^\top QU) = \text{tr}(QUU^\top) \leq \sum_{i=1}^n \lambda_i \sigma_i^2,
\]

where \( \lambda_1 \geq \lambda_2 \geq \ldots \geq \lambda_n \) are eigenvalues of \( Q \) and \( \sigma_1 \geq \sigma_2 \geq \ldots \geq \sigma_n > 0 \) are singular values of \( U \). The equality holds for (4.1) when \( U = V^\top \text{diag}(\sigma_1, \sigma_2, \ldots, \sigma_n)V \),
where the eigenvalue decomposition of $Q$ is given by $Q = V^T \text{diag}(\lambda_1, \lambda_2, \ldots, \lambda_n)V$. Thus, if $\sigma_1$ and $\sigma_n$ are fixed, the maximum value of $\text{tr} Q'$ is given by

$$
\max_{\sigma_2, \ldots, \sigma_{n-1}} \text{tr} Q' = \sum_{i=1}^l \lambda_i \sigma_1^2 + \sum_{i=l+1}^n \lambda_i \sigma_n^2,
$$

where $l$ is the index determined by $\lambda_1 \geq \ldots \geq \lambda_l \geq 0 > \lambda_{l+1} \geq \ldots \geq \lambda_n$. Therefore, if $Q$ has at least one positive eigenvalue, we can generate $Q'$ so as to satisfy $\text{tr} Q' > 0$, that makes it possible to use the same error-bound analysis to $P'_T$.

Based on the above discussion, we consider the case where the form of $U$ is given as $U = V^T \text{diag}(\sigma_1, \ldots, \sigma_1, \sigma_n, \ldots, \sigma_n)V$ ($l \sigma_1$ and $n-l \sigma_n$).

**Lemma 4.2.** There exists a closed ball $B \left(0, \frac{r}{\|U\|}\right)$ which is contained in the polytope $A'z \leq b - Ax_0$.

**Proof.**

If $\|z\| \leq \frac{r}{\|U\|}$, then $\|Uz\| \leq \|U\| \|z\| \leq r$. Since $B(0, r) \subset \{x \mid Ax \leq b - Ax_0\}$, we have $Uz \in \{x \mid Ax \leq b - Ax_0\}$ and $A'z = A(Uz) \leq b - Ax_0$.

**Theorem 4.3.** (Approximation theorem under Assumption 4.1). Define $\tilde{r} \equiv \frac{\|Q\|_F^2}{\|Q\|^2}$. Let $\varepsilon, \varepsilon_1, \varepsilon_3, \varepsilon_4 \in (0, 1], \delta_1, \delta_2 \in (0, 1/2)$.

Suppose that

(i) $d \geq \max \left\{ \frac{\log((24 \text{rank} Q + 6)/\delta_1)}{C_0 \varepsilon_1^2}, \frac{\log(6/\delta_1)}{C_0 \varepsilon_4^2}, \frac{\log(12m/\delta_1)}{C_0 \varepsilon_2^2} \right\}$,

(ii) $d \geq \frac{12C^6}{C_1 \varepsilon_3^4} - \frac{\log \delta_2}{3}$,

(iii) $d < \frac{2C^3 \tilde{r} - \text{cond}(U)^4 \tilde{\delta}}{\varepsilon_3 \text{cond}(U)^2 \sqrt{d}}$.

Then with probability at least $1 - \delta_1 - \delta_2$, the following inequality holds

$$
(4.2) \quad \text{opt}(\text{CRP}_T) \geq \text{opt}(P_T) = \text{opt}(P'_T) \geq \left(1 + \varepsilon \text{cond}(U) \frac{\|x^*\|}{\|y^*\|} \right) \text{opt}(\text{CRP}'_T) - (3\varepsilon_1 + 2\varepsilon_3) \text{cond}(U)^2 \|y^*\|^2 \|Q\|_F,
$$

$$
- \varepsilon_4 \text{cond}(U) \|x^*\| \|2Qx_0 + c\|.
$$

**Proof.** Define $\tilde{r}' = \frac{\|Q'\|_F^2}{\|Q'\|^2}, \tilde{\delta}' = \frac{\text{tr} Q'}{\|Q'\|}$, $\Lambda = \text{diag}(\lambda_1, \lambda_2, \ldots, \lambda_n)$, and $\Sigma = \text{diag}(\sigma_1, \sigma_2, \ldots, \sigma_n)$. First, we observe that

$$
\|Q\| = \|U^T QU\| = \|V^T \Sigma \Lambda V\| = \|\Sigma \Lambda \Sigma\| \leq \sigma_1^2 \|\Lambda\| = \sigma_1^2 \|Q\|,
$$

$$
\|Q\|_F = \|U^T QU\|_F = \|V^T \Sigma \Lambda V\|_F = \|\Sigma \Lambda \Sigma\|_F \geq \sigma_n^2 \|\Lambda\|_F = \sigma_n^2 \|Q\|_F,
$$

$$
\text{tr} Q' = \text{tr}(U^T QU) = \text{tr}(V^T \Sigma \Lambda V) = \text{tr}(\Sigma \Lambda) \leq \sigma_1^2 \text{tr} \Lambda = \sigma_1^2 \text{tr} Q.
$$
Then, on the condition (iii), we have
\[
\frac{2C^3\varepsilon' - \hat{k}'}{\varepsilon_3\sqrt{T'}} = \frac{1}{\varepsilon_3} \left( 2C^3 \frac{\|Q\|_F}{\|Q'\|} - \frac{\text{tr} Q'}{\|Q'\|_F} \right)
\geq \frac{1}{\varepsilon_3} \left( 2C^3 \frac{\|Q\|_F}{\text{cond}(U)^2 \|Q\|} - \text{cond}(U)^2 \frac{\text{tr} Q}{\|Q\|_F} \right)
= \frac{1}{\varepsilon_3 \text{cond}(U)^2} \left( 2C^3 \sqrt{F} - \text{cond}(U)^4 \hat{k} \right)
= \frac{2C^3\varepsilon' - \text{cond}(U)^4 \hat{k}}{\varepsilon_3 \text{cond}(U)^2 \sqrt{F}}.
\]

On the other hand, for the condition (ii), it is easy to see that
\[
\frac{12C^6}{C_1^2\varepsilon_3^3} - \log \delta_2 - \frac{2}{3} \frac{\hat{k}'}{\varepsilon_3 \sqrt{F}} \geq -\frac{2}{3} \frac{\hat{k}'}{\varepsilon_3 \sqrt{F}} + \frac{12C^6}{C_1^2\varepsilon_3^3} - \log \delta_2.
\]
Thus, under the conditions (i)-(iii), we have
(i) \[d \geq \max \left\{ \log((24 \text{ rank } Q + 6)/\delta_1), \log(6/\delta_1), \log(12n/\delta_1) \right\}, \]
(ii) \[d \geq -\frac{2}{3} \frac{\hat{k}'}{\varepsilon_3 \sqrt{F}} + \frac{12C^6}{C_1^2\varepsilon_3^3} - \log \delta_2, \]
(iii) \[d < -\frac{2C^3\varepsilon' - \hat{k}'}{\varepsilon_3 \sqrt{F}}, \]
which are the same to conditions in Theorem 3.14. By Theorem 3.14 and Lemma 4.2, we have, with probability at least 1 - \(\delta_1 - \delta_2\),
\[
\text{opt}(\text{CRP}'_T) \geq \text{opt}(P'_T) \geq \left(1 + \frac{\varepsilon_1}{\varepsilon_3}\right) \frac{\text{opt}(\text{CRP}'_T)}{\varepsilon_3^2} - (3\varepsilon_1 + 2\varepsilon_3) \frac{\|z^*\|^2 \|Q'\|_F}{\|z^*\|} - \varepsilon_1 \frac{\|z^*\| \|U^T(2Qx_0 + c)\|}{\varepsilon_3^2}.
\]
Note that \(\|z^*\| = \|U^{-1}y^*\| \leq \|U^{-1}\| \|y^*\|\), which leads to
\[
\frac{\|z^*\|}{\frac{\varepsilon_3^2}{\|U\|}} \leq \|U\| \|U^{-1}\| \frac{\|y^*\|}{\varepsilon_3} = \text{cond}(U) \frac{\|y^*\|}{\varepsilon_3}, \\
\frac{\|z^*\|^2 \|Q'\|_F}{\|z^*\| \|U^T(2Qx_0 + c)\|} = \|U^{-1}\| \|U\| \frac{\|y^*\|^2 \|Q\|_F}{\|y^*\| \|U^T(2Qx_0 + c)\|} = \text{cond}(U) \frac{\|y^*\|^2 \|Q\|_F}{\|y^*\| \|U^T(2Qx_0 + c)\|}.
\]
By using these inequalities into (4.3), we obtain an error bound in the claim.

We can derive an approximation error of CRP' in a multiplicative form similar to Corollary 3.16 under Assumption 4.1, though we omit the description.

5. Numerical Experiments.

5.1. Randomly generated problems. We perform some experiments on randomly generated non-convex QPs. In the previous sections, we discussed the error between the optimal values of the original problem P and of the convexified projected problem CRP. In this section, to estimate the error in practice, we compare
opt(P), opt(RP) and opt(CRP) so that the errors by random projection and convexification can be verified separately.

Unfortunately, it is difficult to find global optimal solutions of $P$ and $RP$ because they are non-convex QPs. Therefore, we use D.C. algorithms [19] with a multi-start strategy with 10 randomly chosen initial points to find a best possible approximated optimal value. Thus, in this section, opt(P) and opt(RP) denote best possible approximation values of the true optimal values.

Random instances are generated as follows: $Q$ is a diagonal matrix whose diagonal entries are first drawn from some distribution independently and next normalized to $\|Q\|_F = 1$. $c = 1/\sqrt{n1}$, $A_i$ ($i = 1, 2, \ldots, m'$) are random unit vectors and $b_i = 1$. We also add the constraints $-1 \leq x \leq 1$ to ensure the boundedness of the feasible region. The total number of constraints is given by $m = m' + 2n$.

We set $n = 200$, $m' = 5000$, $d = 90, 120, 150, 180$ and the distribution for randomly chosen diagonal entries of $Q$ is $N(\mu, 1^2)$, where $\mu$ is the parameter relating to the convexity of the original problem. More precisely, the larger $\mu$ is, the more positively the eigenvalue distribution of $Q$ is skewed. We calculated optimal values 10 times each for fixed $(d, \mu)$. The average and the standard deviation of the difference between opt(P), opt(RP) and opt(CRP) are shown in Figure 2.

Figure 2(a) shows opt(RP) − opt(P), (b) opt(CRP) − opt(RP) and (c) opt(CRP) − opt(P) versus $\mu$.

We set $n = 200$, $m' = 5000$, $d = 90, 120, 150, 180$ and the distribution for randomly chosen diagonal entries of $Q$ is $N(\mu, 1^2)$, where $\mu$ is the parameter relating to the convexity of the original problem. More precisely, the larger $\mu$ is, the more positively the eigenvalue distribution of $Q$ is skewed. We calculated optimal values 10 times each for fixed $(d, \mu)$. The average and the standard deviation of the difference between opt(P), opt(RP) and opt(CRP) are shown in Figure 2.

Figure 2(a) shows opt(RP) − opt(P), the error due to random projection. We
confirm that \( \text{opt}(\text{RP}) - \text{opt}(\text{P}) \) gets smaller as \( d \) gets larger. This is because as \( d \) gets larger, random projections become more likely to preserve geometric quantities or function values \( \text{(Lemma 2.11)} \).

Next we discuss the results shown in Figure 2(b), that is the error due to convexification: \( \text{opt}(\text{CRP}) - \text{opt}(\text{RP}) \). The first observation we can make is that \( d \) should be smaller for \( \text{opt}(\text{CRP}) - \text{opt}(\text{RP}) \) to be smaller. This is the opposite of the previous observation. This fact comes from \( \text{Lemma 3.1} \). Indeed, for \( \text{opt}(\text{CRP}) - \text{opt}(\text{RP}) \) to be small, \( Q^+ \) must be a good approximation of \( Q \), or equivalently, most eigenvalues of \( Q \) must be positive, which will be satisfied by setting \( d \) small since \( Q \approx \frac{mt}{d}I_d \). We also see that \( \text{opt}(\text{CRP}) - \text{opt}(\text{RP}) \) decreases monotonically with respect to \( \mu \). This is because if \( \mu \) is large, then \( \text{P} \) and \( \text{RP} \) will be nearly convex problems and the error caused by convexification will be small.

Figure 2(c) shows \( \text{opt}(\text{CRP}) - \text{opt}(\text{P}) \) we mainly discuss in this paper. Since we use convexification, the error between \( \text{opt}(\text{CRP}) \) and \( \text{opt}(\text{P}) \) highly depends on the eigenvalue distribution of \( Q \). Our method behaves better when the percentage of positive eigenvalues of \( Q \) is large.

5.2. Application for Support Vector Machine Classification with Indefinite Kernels. Let \( K \in \mathbb{R}^{d \times n} \) be a given kernel matrix and \( y \in \mathbb{R}^n \) be the vector of labels, with \( Y = \text{diag}(y) \). The classic soft margin SVM problem \( \text{[4, 11]} \) is formulated as:

\[
(5.1) \quad \min_{\alpha} \{ \alpha^T YKY\alpha - 2\alpha^T 1 \mid 0 \leq \alpha \leq C1, \alpha^Ty = 0 \},
\]

where \( \alpha \in \mathbb{R}^n \) and \( C \) is the SVM misclassification penalty which is fixed to 1 in this paper. There are some works (see e.g., \[17, 8, 14, 12] \) investigating applications where kernel matrices formed using similarity measures are not positive semidefinite and algorithms for SVM \( \text{(5.1)} \) with indefinite \( K \). If \( K \) is an indefinite kernel matrix, \( \text{(5.1)} \) is a non-convex QP. The corresponding convexified randomly projected problem is

\[
(5.2) \quad \min_{u} \{ u^TPYKYP^Tu - 2u^TP1 \mid 0 \leq P^Tu \leq C1, u^TPy = 0 \},
\]

where \( P \in \mathbb{R}^{d \times n} \) is a random matrix and \( u \in \mathbb{R}^d \). Although we can not apply our theoretical guarantees because the original problem does not have a full dimensional feasible region, we expect \( P^Tu^* \) to be a good approximation of the optimum of the problem \( \text{(5.1)} \), where \( u^* \) is an optimum of the convexified randomly projected problem \( \text{(5.2)} \).

We performed the experiments on the image data of 0, 1 and 7 from the MNIST handwritten digits database \[13] using the indefinite simpson score \[12] as a kernel function value to measure the similarity of two images. We experimented with three different binary classifications: 0 and 1, 0 and 7, and 1 and 7. In all cases, we choose 1000\((= n)\) train data where each class has 500 points or images and 400 test data where each class has 200. The results are shown in Tables 2 to 4. We have solved \( \text{(5.2)} \) 20 times with different random \( P \) for each \( d \) and evaluated the optimum of \( \text{(5.2)} \) with test data. “Training Accuracy” and “Test Accuracy” in the tables refer to the average and standard deviation among 20 training-accuracy and test-accuracy values. We confirmed that SVM with simpson score works to find a good approximate solution of the original problem \( \text{(5.1)} \) for appropriate \( d \). We also calculated the accuracy using the optimal solution of the following problem obtained by convexifying \( \text{(5.1)} \) directly:

\[
(5.3) \quad \min_{\alpha} \{ \alpha^T YKY\alpha - 2\alpha^T 1 \mid 0 \leq \alpha \leq C1, \alpha^Ty = 0 \},
\]
and obtained 63.90% training-accuracy and 68.50% test-accuracy, 62.00% training-accuracy and 61.00% test-accuracy and 66.70% training-accuracy and 53.00% test-accuracy for the binary classification of 0 and 1, 0 and 7, and 1 and 7, respectively, so that we conclude that combining random projections and convexification performs as well or better than just convexification alone.

6. Conclusions. Random projections have been applied to solve optimization problems in suitable lower-dimensional spaces in various existing works. However, to the best of our knowledge, it is the first time they are used to build a convex approximation for a non-convex quadratic optimization problem. In this paper, we proved that the randomly projected problem $\mathbf{RP}$ that is proposed in [5] is close to a convex problem. This allowed us to propose a convexified randomly projected problem, $\mathbf{CRP}$, that we used to obtain an approximate optimal value of $\mathbf{P}$.

In our framework, the existence of a value $d$, that will correspond to the dimension after projections, depends on the distribution of the eigenvalues of $\mathbf{Q}$. We proved that even if $\text{tr} \mathbf{Q}$ is negative then, under some additional error cost, we could use scaling and preconditioning to transform the problem into a new one where the theory applies. To confirm that our method is practical, we applied our framework to SVM classification problem with indefinite kernel, though the problem setting does not satisfy the conditions necessary for the theoretical guarantee. As shown in subsection 5.2, our method is able to find good approximate global optimal solutions by only solving $\mathbf{CRP}$, which scores as well or better than solving a problem that is only a convexification of the original problem. At least, it is worth trying our method for a non-convex quadratic problem since $\mathbf{CRP}$ is convex and its size is smaller than the original problem and $\text{opt}(\mathbf{CRP})$ can be obtained by the solver with few computational resources.

One of the directions for future research is to generalize the objective function and constraints, which is still difficult since our argument depends on Lemma 2.11 that shows that random projections preserve linear or quadratic function values. For a general objective function, we can consider an iterative method using quadratic approximation of the function at each point, but obtaining theoretical guarantees in such a case needs further investigations.

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Table 2
Accuracy (average and standard deviation of 20 times for each $d$) of (5.2) for the MNIST 0-1, while (5.3) achieved 63.90% training-accuracy and 68.50% test-accuracy

| $d$ | Training Accuracy (%) | Test Accuracy (%) |
|-----|-----------------------|-------------------|
| 300 | 48.47 ± 17.13         | 48.45 ± 16.89     |
| 400 | 62.92 ± 17.90         | 63.55 ± 20.13     |
| 500 | 80.16 ± 17.17         | 82.25 ± 18.53     |
| 600 | 95.58 ± 0.43          | 97.18 ± 0.40      |
| 700 | 97.12 ± 0.27          | 98.48 ± 0.46      |
| 800 | 97.15 ± 0.56          | 96.53 ± 0.43      |
| 900 | 88.40 ± 3.01          | 87.53 ± 4.12      |
| 1000| 64.60 ± 1.56          | 68.53 ± 2.40      |

Table 3
Accuracy (average and standard deviation of 20 times for each $d$) of (5.2) for the MNIST 0-7, while (5.3) achieved 62.00% training-accuracy and 61.00% test-accuracy

| $d$ | Training Accuracy (%) | Test Accuracy (%) |
|-----|-----------------------|-------------------|
| 300 | 49.90 ± 9.24          | 49.40 ± 10.27     |
| 400 | 59.67 ± 14.56         | 61.83 ± 16.52     |
| 500 | 76.99 ± 20.22         | 79.18 ± 20.56     |
| 600 | 95.73 ± 0.29          | 98.43 ± 0.18      |
| 700 | 94.25 ± 0.85          | 95.53 ± 0.91      |
| 800 | 84.00 ± 2.99          | 81.93 ± 3.95      |
| 900 | 69.37 ± 3.42          | 66.15 ± 2.57      |
| 1000| 62.42 ± 0.96          | 61.05 ± 0.57      |

Table 4
Accuracy (average and standard deviation of 20 times for each $d$) of (5.2) for the MNIST 1-7, while (5.3) achieved 66.70% training-accuracy and 53.00% test-accuracy

| $d$ | Training Accuracy (%) | Test Accuracy (%) |
|-----|-----------------------|-------------------|
| 300 | 55.09 ± 14.40         | 53.43 ± 12.80     |
| 400 | 64.88 ± 15.64         | 62.80 ± 14.14     |
| 500 | 89.00 ± 12.61         | 83.30 ± 11.56     |
| 600 | 95.00 ± 0.85          | 90.23 ± 1.54      |
| 700 | 92.56 ± 1.60          | 86.53 ± 2.95      |
| 800 | 84.26 ± 3.09          | 76.60 ± 4.27      |
| 900 | 72.44 ± 3.20          | 60.83 ± 4.19      |
| 1000| 66.96 ± 0.63          | 53.60 ± 0.78      |