Absence of Replica Symmetry Breaking in the Random Field Ising Model

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Abstract: It is shown that replica symmetry is not broken in the random field Ising model in any dimension, at any temperature and field strength, except possibly at a measure-zero set of exceptional temperatures and field strengths.

1. Introduction

Let \( \mathbb{Z}^d \) denote the \( d \)-dimensional integer lattice. For each \( n \), let \( V_n := \mathbb{Z}^d \cap [1, n]^d \). The random (Gaussian) field Ising model (RFIM) on \( V_n \) with free boundary condition is the random probability measure on the set of spin configurations \( \{-1, 1\}^{V_n} \) that is proportional to

\[
\exp \left( \beta \sum_{\langle xy \rangle} \sigma_x \sigma_y + h \sum_x g_x \sigma_x \right),
\]

where \( \langle xy \rangle \) below the first sum means that we are summing over \( x, y \) that are neighbors, \( \beta \) and \( h \) are positive numbers that we will call inverse temperature and field strength, and \( g_x \) are independent standard Gaussian random variables. The variables \( g_x \) are collectively called the disorder.

The RFIM was introduced in a seminal paper of Imry and Ma [11] as a simple example of a disordered system. The model is often showcased as a notable success story of mathematical physics. In the paper [11], Imry and Ma gave an intuitive argument that the model does not have an ordered phase in dimensions one and two, but does in dimensions three and higher. There was spirited controversy in the theoretical physics community about this claim for more than 10 years after the introduction of this model, with arguments coming both for and against. The question was finally settled in two rigorous papers by Bricmont and Kupiainen [5,6], who proved that there is indeed an

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ordered phase in dimensions three and higher, and two other rigorous papers by Aizen-
mans and Wehr [2,3], who proved the nonexistence of an ordered phase in dimension
two. (The one dimensional case was settled earlier.) For a readable account of these
proofs and an up-to-date survey of the literature, see [4, Chapter 7].

However, in spite of all this progress, the RFIM has not yet yielded an exact solu-
tion, and there are many aspects that remain unknown. A popular theoretical physics
approach to solving statistical models of disordered systems is the so-called ‘replica
method’. There have been a number of attempts to apply this technique to the RFIM
by theoretical physicists (e.g. in [15]), with varying degrees of success. A crucial step
in the replica approach is to understand whether the model under investigation exhibits
‘replica symmetry breaking’. The purpose of this paper is to investigate this question for
the RFIM and rigorously prove that replica symmetry does not break in this model. This
lends support to the recent finding in the physics literature that the RFIM does not have
a spin glass phase [12,13] and provides partial justification for the replica symmetric
solution of this model.

It is worth stressing, however, that even from the point of view of the physicists, the
replica symmetric solution to the RFIM is not the final answer. There is a general claim,
not yet verified rigorously, that the RFIM has a large number (increasing with the system
size) of so-called ‘metastable’ states. These metastable states play an important role, but
how to identify and count them is still not clear (other than in some simple cases, see
e.g. the case of RFIM on random graphs discussed in [12]).

Let me now briefly explain what is meant by the absence of replica symmetry break-
ing. Given a fixed realization of the disorder, the probability measure defined above is
called the Gibbs measure of the system. Suppose that two spin configurations
\( \sigma_1 \) and \( \sigma_2 \) are drawn independently from this random Gibbs measure. These are called replicas.
The overlap between these two replicas is defined as

\[
R_{1,2} := \frac{1}{|V_n|} \sum_x \sigma_x^1 \sigma_x^2.
\]

The system is said to exhibit replica symmetry breaking (e.g., according to Parisi [17])
if the limiting distribution of the random variable \( R_{1,2} \), when \( n \to \infty \), has more than
one point in its support. The following theorem shows that this does not happen for the
RFIM.

**Theorem 1.1.** There is a set \( A \subseteq (0, \infty)^2 \) of Lebesgue measure zero such that for any
\((\beta, h) \in A\), there exists a constant \( q_{\beta, h} \in [-1, 1] \) such that

\[
\lim_{n \to \infty} \mathbb{E}\langle (R_{1,2} - q_{\beta, h})^2 \rangle = 0,
\]

where \( \langle \cdot \rangle \) denotes averaging with respect to the Gibbs measure and \( \mathbb{E} \) is expectation with
respect to the disorder. Moreover, for every \( \beta \), the set of all \( h \) such that \((\beta, h) \in A \) is
countable.

The rest of this paper is devoted to the proof of the above theorem. The proof
depends on two key ingredients: the FKG property of the RFIM, and the Ghirlanda–
Guerra identities. The proof should extend easily to any other model that has these two
properties.
2. Proof

For every $\beta$, $h$ and $n$, define

$$Z_n(\beta, h) := \sum_{\sigma \in \{-1, 1\}^{V_n}} \exp \left( \beta \sum_{(xy)} \sigma_x \sigma_y + h \sum_x g_x \sigma_x \right).$$

Let $F_n(\beta, h) := \log Z_n(\beta, h)$. Let

$$\psi_n(\beta, h) := \frac{F_n(\beta, h)}{|V_n|}$$

and $p_n(\beta, h) := \mathbb{E}(\psi_n(\beta, h))$.

Lemma 2.1. For every positive $\beta$ and $h$, the limit

$$p(\beta, h) := \lim_{n \to \infty} p_n(\beta, h)$$

exists and is finite.

The proof of this lemma is quite standard, but may be hard to find. For the sake of completeness, a proof is given below.

Proof. Fix $\beta$ and $h$. Let $Z_n = Z_n(\beta, h)$, $F_n = F_n(\beta, h)$ and $p_n = p_n(\beta, h)$. Take any two integers $m$ and $n$, and let $l := mn$. Then the vertices of the box $V_l$ may be partitioned into $m^d$ disjoint translates of $V_n$ that we will call ‘blocks’. Define a new Hamiltonian on $V_l$ by taking the old one and deleting the terms $\sigma_x \sigma_y$ when $x$ and $y$ belong to different blocks. Let $Z'_l$ be the partition function for this new Hamiltonian. Then clearly $Z'_l/Z_l$ is bounded above and below by $e^{\beta L}$ and $e^{-\beta L}$ respectively, where $L$ is the number of deleted bonds. Note that $L \leq C(d)n^{d-1}m^d$, where $C(d)$ is a constant that depends on $d$ only. Thus,

$$|\log Z'_l - \log Z_l| \leq C(d)\beta n^{d-1}m^d.$$  

But $Z'_l$ factorizes into $m^d$ independent terms, each of which has the same distribution as $Z_n$. Thus, $\mathbb{E}(\log Z'_l) = m^d\mathbb{E}(F_n)$. By the above inequality, this shows that

$$|p_n - p_l| = |n^{-d}\mathbb{E}(F_n) - l^{-d}\mathbb{E}(F_l)| 
= l^{-d}|m^d\mathbb{E}(F_n) - \mathbb{E}(F_l)| 
\leq l^{-d}C(d)\beta n^{d-1}m^d = C(d)\beta n^{-1}. $$

By the symmetry of the situation, $|p_m - p_l| \leq C(d)\beta m^{-1}$. Combining, we get $|p_m - p_n| \leq C(d)\beta(m^{-1} + n^{-1})$. This shows that the sequence $p_n$ is Cauchy and hence convergent. To show that the limit is finite, simply apply Jensen’s inequality while trying to evaluate the expectation in $p_n = n^{-d}\mathbb{E}(F_n)$. \qed

The next lemma is also standard.

Lemma 2.2. The limit $p(\beta, h)$ defined in Lemma 2.1 is a convex function of $h$ for every fixed $\beta$. The same is true for $F_n$, $\psi_n$ and $p_n$.

Proof. It is easy to differentiate $F_n$ twice in $h$ and show that the derivative is positive, which implies convexity of $F_n$ as a function of $h$. This shows that $p_n$, and therefore $p$, is convex. \qed
Let $A$ be the set of all $(\beta, h) \in (0, \infty)^2$ such that the function $p$ is not differentiable in $h$ at the point $(\beta, h)$.

**Lemma 2.3.** The set $A$ has Lebesgue measure zero. Moreover, for every $\beta$, the set of all $h$ such that $(\beta, h) \in A$ is countable.

**Proof.** For fixed $\beta$, $p$ is a convex function of $h$ by Lemma 2.2. It is well known (and easy to prove) that the set of points at which a convex function on $(0, \infty)$ is not differentiable must be countable. This proves the second claim of the lemma. For the first, use the second, together with a simple application of Fubini’s theorem. □

The following lemma is also quite standard; as before, a proof is provided for the sake of completeness.

**Lemma 2.4.** For any $\beta$, $h$ and $n$,

$$\text{Var}(F_n(\beta, h)) \leq h^2 |V_n|.$$  

**Proof.** Note that by the Poincaré inequality for the Gaussian measure [14, p. 49],

$$\text{Var}(F_n(\beta, h)) \leq \sum_{x \in V_n} \mathbb{E}\left(\left(\frac{\partial F_n(\beta, h)}{\partial g_x}\right)^2\right) = h^2 \sum_{x \in V_n} \mathbb{E}(\sigma_x)^2$$

Since $(\sigma_x)^2 \leq 1$ for all $x$, this completes the proof of the lemma. □

The next lemma records that the RFIM has the FKG property.

**Lemma 2.5.** Let $f$ and $g$ be two monotone increasing functions on the configuration space $\{-1, 1\}^{V_n}$. Then $\langle fg \rangle \geq \langle f \rangle \langle g \rangle$. In particular, for any $x$ and $y$, $\langle \sigma_x \sigma_y \rangle \geq \langle \sigma_x \rangle \langle \sigma_y \rangle$.

**Proof.** It is simple to verify that for any value of the disorder, the Gibbs measure of the RFIM satisfies the FKG lattice condition [8], which implies positive correlations between increasing functions. □

The next lemma is the key ingredient in the proof of Theorem 1.1 and the main original component of this paper. The proof makes a crucial use of the FKG property of the RFIM.

**Lemma 2.6.** For any $n$ and any $(\beta, h) \in (0, \infty)^2$,

$$\mathbb{E}\left(\langle R_{1,2}^2 \rangle - \langle R_{1,2} \rangle^2\right) \leq \frac{2\sqrt{2 + h^2}}{h \sqrt{|V_n|}}.$$  

**Proof.** Fix $\beta$, $h$ and $n$, and let $F := F_n(\beta, h) - \mathbb{E}(F_n(\beta, h))$. Let $W$ denote the set of all unordered pairs $\{x, y\}$, where $x, y \in V_n$ and $x \neq y$. For each $w = \{x, y\} \in W$, let $s_w := g_x g_y$, and let

$$c_w := \mathbb{E}(s_w F).$$

Define

$$S := \sum_{w \in W} c_w s_w.$$
Note that for any two distinct $w, w' \in W$, $\mathbb{E}(s_w s_{w'}) = 0$, and for any $w$, $\mathbb{E}(s_w^2) = 1$. Thus,
\[
\mathbb{E}(S^2) = \sum_w c_w^2 = \sum_w \mathbb{E}(c_w s_w F) = \mathbb{E}(SF).
\]
Consequently,
\[
\mathbb{E}((F - S)^2) = \mathbb{E}(F^2) - 2\mathbb{E}(FS) + \mathbb{E}(S^2) = \mathbb{E}(F^2) - \mathbb{E}(S^2).
\]
Combining the last two displays and applying Lemma 2.4 gives
\[
\sum_w c_w^2 \leq \mathbb{E}(F^2) = \text{Var}(F_n(\beta, h)) \leq h^2|V_n|.
\]
Applying integration by parts, for any $w = \{x, y\}$,
\[
c_w = \mathbb{E}(g_x g_y F) = \mathbb{E}\left(\frac{\partial^2 F}{\partial g_x \partial g_y}\right) = h^2 \mathbb{E}(\langle \sigma_x \sigma_y \rangle - \langle \sigma_x \rangle \langle \sigma_y \rangle) = : h^2 r_{x,y}.
\]
Therefore by (1),
\[
\sum_{x,y \in V_n} r_{x,y}^2 \leq 2 \sum_{\{x,y\} \in W} r_{x,y}^2 + |V_n| \leq \frac{(2 + h^2)|V_n|}{h^2}.
\]
Now note that
\[
\mathbb{E}((R_{1,2}^2) - (R_{1,2})^2) = \frac{1}{|V_n|^2} \sum_{x,y} \mathbb{E}((\langle \sigma_x \sigma_y \rangle - \langle \sigma_x \rangle \langle \sigma_y \rangle)^2)
\]
\[
\leq \frac{1}{|V_n|^2} \sum_{x,y} \mathbb{E}|\langle \sigma_x \sigma_y \rangle - \langle \sigma_x \rangle \langle \sigma_y \rangle| \langle \sigma_x \sigma_y \rangle + \langle \sigma_x \rangle \langle \sigma_y \rangle|
\]
\[
\leq \frac{2}{|V_n|^2} \sum_{x,y} \mathbb{E}|\langle \sigma_x \sigma_y \rangle - \langle \sigma_x \rangle \langle \sigma_y \rangle|.
\]
But by the FKG property of the RFIM (Lemma 2.5), $\langle \sigma_x \sigma_y \rangle - \langle \sigma_x \rangle \langle \sigma_y \rangle \geq 0$ for each $x$ and $y$. Thus, by the above inequality and (2),
\[
\mathbb{E}((R_{1,2}^2) - (R_{1,2})^2) \leq \frac{2}{|V_n|^2} \sum_{x,y} r_{x,y}
\]
\[
\leq \frac{2}{|V_n|^2} \left(|V_n|^2 \sum_{x,y} r_{x,y}^2\right)^{1/2} \leq \frac{2\sqrt{2} + h^2}{h\sqrt{|V_n|}}.
\]
This completes the proof of the lemma. \(\square\)

Next, define
\[
H_n := \frac{1}{|V_n|} \sum_{x \in V_n} g_x \sigma_x.
\]
Recall the set $A$ defined in the paragraph preceding Lemma 2.3. Let $A^c$ denote the complement of $A$ in $\mathbb{R}^2$. 

Lemma 2.7. For any \((\beta, h) \in A^c\),

\[
\lim_{n \to \infty} E(H_n) = \frac{\partial p}{\partial h}(\beta, h)
\]

and

\[
\lim_{n \to \infty} E|\langle H_n \rangle - E(H_n)| = 0.
\]

Proof. Note that

\[
\langle H_n \rangle = \frac{\partial \psi_n}{\partial h} \quad \text{and} \quad E\langle H_n \rangle = \frac{\partial p_n}{\partial h}.
\]

Fix some \(h' > h > 0\). By the convexity of \(\psi_n\) (Lemma 2.2) and the first identity in the above display,

\[
\langle H_n \rangle \leq \frac{\psi_n(\beta, h') - \psi_n(\beta, h)}{h' - h}.
\]

By Lemma 2.4,

\[
E|\psi_n(\beta, h) - p_n(\beta, h)| \leq \sqrt{\text{Var}(\psi_n(\beta, h))} \leq \frac{h}{\sqrt{|V_n|}}.
\]

Therefore, if \(p'\) denotes the function \(\partial p/\partial h\), then

\[
E \left| \frac{\psi_n(\beta, h') - \psi_n(\beta, h)}{h' - h} - p'(\beta, h) \right|
\leq \frac{h + h'}{h' - h} \frac{1}{\sqrt{|V_n|}} \left| \frac{|p_n(\beta, h') - p(\beta, h')| + |p_n(\beta, h) - p(\beta, h)|}{h' - h} + \frac{|p(\beta, h') - p(\beta, h)|}{h' - h} - p'(\beta, h) \right|.
\]

Since \(|V_n| \to \infty\) and \(p_n(\beta, h) \to p(\beta, h)\) by Lemma 2.1, this gives

\[
\limsup_{n \to \infty} E \left| \frac{\psi_n(\beta, h') - \psi_n(\beta, h)}{h' - h} - p'(\beta, h) \right| \leq \frac{|p(\beta, h') - p(\beta, h)|}{h' - h} - p'(\beta, h).
\]

Combining with (3), we get

\[
\limsup_{n \to \infty} E((H_n) - p'(\beta, h))^+ \leq \frac{|p(\beta, h') - p(\beta, h)|}{h' - h} - p'(\beta, h),
\]

where \(x^+\) denotes the positive part of a real number \(x\). Since this bound holds for any \(h' > h\) and \(p\) is differentiable at \(h\),

\[
\lim_{n \to \infty} E((H_n) - p'(\beta, h))^+ = 0.
\]

Similarly, considering \(h' < h\) and repeating the steps, we can show that the limit of the negative part is zero as well. Thus,

\[
\lim_{n \to \infty} E|\langle H_n \rangle - p'(\beta, h)| = 0.
\]
By Jensen’s inequality, this gives
\[\lim_{n \to \infty} |\mathbb{E}\langle H_n \rangle - p'(\beta, h)| = 0.\]

The proof is now easily completed by combining the last two inequalities. □

**Lemma 2.8.** For any \(\beta, h\) and \(n\),
\[
\sum_{x, y} \left( \mathbb{E}\left( \frac{\partial^4 F_n}{\partial g_x^2 \partial g_y^2} \right) \right)^2 \leq 24h^2|V_n|.
\]

*Proof.* This proof is very similar to that of Lemma 2.6. Let \(F := F_n(\beta, h) - \mathbb{E}(F_n(\beta, h))\). Let \(U\) be the set of all unordered pairs \(\{x, y\}\), where \(x, y \in V_n\). For any \(u = \{x, y\} \in U\), define
\[t_u := \begin{cases} 
\frac{1}{2}(g_x^2 - 1)(g_y^2 - 1) & \text{if } x \neq y, \\
\frac{1}{\sqrt{24}}(g_x^4 - 6g_x^2 + 3) & \text{if } x = y.
\end{cases}
\]
Define \(d_u := \mathbb{E}(t_u F)\) and
\[T := \sum_{u \in U} d_u t_u.
\]
An easy computation shows that for any two distinct \(u, u' \in U\), \(\mathbb{E}(t_u t_{u'}) = 0\), and for any \(u\), \(\mathbb{E}(t_u^2) = 1\). Thus,
\[\mathbb{E}(T^2) = \sum_u d_u^2 = \sum_u \mathbb{E}(d_u t_u F) = \mathbb{E}(T F).
\]
Consequently,
\[\mathbb{E}(F - T)^2 = \mathbb{E}(F^2) - 2\mathbb{E}(F T) + \mathbb{E}(T^2) = \mathbb{E}(F^2) - \mathbb{E}(T^2).
\]
Combining the last two displays and applying Lemma 2.4 gives
\[\sum_u d_u^2 \leq \mathbb{E}(F^2) = \text{Var}(F_n(\beta, h)) \leq h^2|V_n|.
\]
Now, integration by parts shows that if \(u = \{x, y\}\) for some \(x \neq y\),
\[d_u = \frac{1}{2} \mathbb{E}\left( \frac{\partial^4 F_n}{\partial g_x^2 \partial g_y^2} \right),
\]
and for \(u = \{x, x\}\),
\[d_u = \frac{1}{\sqrt{24}} \mathbb{E}\left( \frac{\partial^4 F_n}{\partial g_x^4} \right).
\]
This completes the proof of the lemma. □

**Lemma 2.9.** For any \((\beta, h) \in A^c\),
\[\lim_{n \to \infty} \mathbb{E}(|H_n - \mathbb{E}(H_n)|) = 0.
\]
Proof. Integration by parts gives

\[
\mathbb{E}(\langle H_n^2 \rangle - \langle H_n \rangle^2) = \frac{1}{|V_n|^2} \sum_{x, y \in V_n} \mathbb{E}(g_x g_y ((\sigma_x \sigma_y) - \langle \sigma_x \sigma_y \rangle))
\]

\[
= \frac{1}{|V_n|^2} \sum_{x, y} \mathbb{E} \left( \frac{\partial^2}{\partial g_x \partial g_y} (\langle \sigma_x \sigma_y \rangle - \langle \sigma_x \rangle \langle \sigma_y \rangle) \right)
\]

\[
+ \frac{1}{|V_n|^2} \sum_x \mathbb{E}(1 - \langle \sigma_x \rangle^2).
\]

Now note that

\[
(\sigma_x \sigma_y) - \langle \sigma_x \rangle \langle \sigma_y \rangle = \frac{1}{h^2} \frac{\partial^2 F_n}{\partial g_x \partial g_y}.
\]

Thus, by the Cauchy–Schwarz inequality and Lemma 2.8,

\[
\mathbb{E}(\langle H_n^2 \rangle - \langle H_n \rangle^2) \leq \frac{1}{h^2 |V_n|^2} \sum_{x, y} \mathbb{E} \left( \frac{\partial^4 F_n}{\partial g_x^2 \partial g_y^2} \right) + \frac{1}{|V_n|}
\]

\[
\leq \frac{1}{h^2 |V_n|^2} \left( |V_n|^2 \sum_{x, y} \mathbb{E} \left( \frac{\partial^4 F_n}{\partial g_x^2 \partial g_y^2} \right)^2 \right)^{1/2} + \frac{1}{|V_n|}
\]

\[
\leq \frac{\sqrt{24}}{h \sqrt{|V_n|}} + \frac{1}{|V_n|}.
\]

Combining this with the inequality

\[
\mathbb{E}(\langle H_n - \langle H_n \rangle \rangle) \leq \sqrt{\mathbb{E}(\langle (H_n - \langle H_n \rangle)^2 \rangle)} = \sqrt{\mathbb{E}(\langle H_n^2 \rangle - \langle H_n \rangle^2)}
\]

implies that \(\mathbb{E}(\langle |H_n - \langle H_n \rangle| \rangle) \to 0\). Since \((\beta, h) \in A^c\), we can now apply Lemma 2.7 to complete the proof. \(\square\)

Take any integer \(k \geq 2\) and let \(\sigma^1, \ldots, \sigma^k, \sigma^{k+1}\) denote \(k + 1\) spin configurations drawn independently from the Gibbs measure. Let \(R_{i, j}\) denote the overlap between \(\sigma^i\) and \(\sigma^j\). Let \(f\) be a function of these overlaps, that takes value in \([-1, 1]\) and does not change with \(n\). The following result shows that the RFIM satisfies the Ghirlanda–Guerra identities [9] (also sometimes called the Aizenman–Contucci identities [1]) at almost all \((\beta, h)\).

Lemma 2.10. If \((\beta, h) \in A^c\), and \(f\) is as above, then

\[
\lim_{n \to \infty} \left( \mathbb{E}(f R_{1,k+1}) - \frac{1}{k} \mathbb{E}(f) \mathbb{E}(R_{1,2}) - \frac{1}{k} \sum_{i=2}^{k} \mathbb{E}(f R_{1,i}) \right) = 0.
\]

Given a result like Lemma 2.9, the derivation of the Ghirlanda–Guerra identities is quite standard [7, 16, 18]. As usual, details are presented below for the sake of completeness.
Proof. Integration by parts gives
\[ E \langle H_n(\sigma^1)f \rangle = \frac{1}{|V_n|} \sum_x E(g_x(\sigma^1_x f)) \]
\[ = \frac{1}{|V_n|} \sum_x E \left( \frac{\partial \langle \sigma^1_x f \rangle}{\partial g_x} \right) \]
\[ = h \sum_{i=1}^k \mathbb{E}(f R_{1,i}) - hk \mathbb{E}(f R_{1,k+1}). \]
In particular,
\[ \mathbb{E}(H_n) = h \mathbb{E}(R_{1,2}). \] (4)
Therefore,
\[ \mathbb{E}(H_n(\sigma^1)f) - \mathbb{E}(H_n)\mathbb{E}(f) \]
\[ = h \mathbb{E}(R_{1,2})\mathbb{E}(f) + h \sum_{i=2}^k \mathbb{E}(f R_{1,i}) - hk \mathbb{E}(f R_{1,k+1}). \]
To complete the proof, note that by Lemma 2.7, the left-hand side in the above display tends to zero as \( n \to \infty \). \( \square \)

Proof of Theorem 1.1. We now have all the ingredients necessary for completing the proof of Theorem 1.1. The main ingredient is Lemma 2.6. This needs to be combined with the Ghirlanda–Guerra identities and a nice trick of Guerra [10] which also appears in Talagrand [18].

Choosing \( k = 2 \) and \( f = R_{1,2} \) in Lemma 2.10 gives
\[ \lim_{n \to \infty} \left( \mathbb{E}(R_{1,2} R_{1,3}) - \frac{1}{2} (\mathbb{E}(R_{1,2}))^2 - \frac{1}{2} \mathbb{E}(R_{1,2}^2) \right) = 0. \] (5)
Choosing \( k = 3 \) and \( f = R_{2,3} \) gives
\[ \lim_{n \to \infty} \left( \mathbb{E}(R_{2,3} R_{1,4}) - \frac{1}{3} (\mathbb{E}(R_{1,2}))^2 - \frac{1}{3} \mathbb{E}(R_{2,3} R_{1,2}) - \frac{1}{3} \mathbb{E}(R_{2,3} R_{1,3}) \right) = 0. \] (6)
Now, since by symmetry
\[ \mathbb{E}(R_{2,3} R_{1,2}) = \mathbb{E}(R_{2,3} R_{1,3}) = \mathbb{E}(R_{1,3} R_{1,2}), \]
we can use (5) to evaluate the last two terms in (6), to get
\[ \lim_{n \to \infty} \left( \mathbb{E}(R_{2,3} R_{1,4}) - \frac{2}{3} (\mathbb{E}(R_{1,2}))^2 - \frac{1}{3} \mathbb{E}(R_{1,2}^2) \right) = 0. \]
Since \( \langle R_{2,3} R_{1,4} \rangle = \langle R_{1,2} \rangle^2 \), we can now combine the above display with Lemma 2.6 to have
\[ \lim_{n \to \infty} (\mathbb{E}(R_{1,2}^2) - (\mathbb{E}(R_{1,2}))^2) = 0. \]
But the term on the left-hand side is nothing but \( \mathbb{E}((R_{1,2} - \mathbb{E}(R_{1,2}))^2) \). To complete the proof, note that by (4) and Lemma 2.7, \( \lim_{n \to \infty} \mathbb{E}(R_{1,2}) \) exists. \( \square \)

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