The regional scale surface mass balance of Pine Island Glacier, West Antarctica over the period 2005–2014, derived from airborne radar soundings and neutron probe measurements

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Abstract. We derive recent surface mass balance (SMB) estimates from airborne radar observations along the iSTAR traverse (2013,2014) at Pine Island Glacier (PIG), West Antarctica. Ground based neutron probe measurements of snow density at 22 locations allow us to derive SMB from the annual internal radar reflection layers. The 2005 layer was traced for a total distance of 2367 km allowing us to determine annual mean SMB for the period 2005–2014. Using complementary SMB estimates from the RACMO2.3p2 regional climate model and a geostatistical kriging scheme we determine a regional scale SMB distribution with the same main characteristics as that determined for the period 1985–2009 in previous studies. Local departures exist for the northern PIG slopes, where the orographic precipitation shadow effect appears to be more pronounced in our observations, and the southward interior, where the SMB gradient is more pronounced in previous studies. For the PIG basin, we derive a total mass input of $79.9 \pm 19.2 \text{ Gt yr}^{-1}$. This is not significantly different to the value of $78.3 \pm 6.8 \text{ Gt yr}^{-1}$ for the period 1985–2009. Thus there is no evidence of a secular trend in mass input to the PIG basin. We note, however, that our estimated uncertainty is more than twice the uncertainty for the 1985–2009 estimates. Our error analysis indicates that uncertainty estimates on total mass input are highly sensitive to the selected kriging methodology and assumptions made on the interpolation error, which we identify as the main cause for the increased uncertainty range compared to the 1985–2009 estimates.

Introduction

The stability of the West Antarctic Ice Sheet (WAIS) is a major concern for scientists seeking to predict global sea level rise. Transport of heat from the upwelling circumpolar deep water has proved to be a critical driver of Antarctic ice shelf thinning and grounding line retreat, thus stimulating the acceleration of marine-terminating outlet glaciers (e.g. Hillenbrand et al., 2017). In particular the Amundsen Sea sector has experienced an unprecedented acceleration in ice discharge since the beginning of satellite based ice flow observations in the 1970s. Three quarters of this ice discharge stem from the Thwaites and Pine Island glaciers with both showing evidence of rapid acceleration since the 1970s (Mouginot et al., 2014) and spreading of surface lowering along their tributaries over the past two decades (Konrad et al., 2017). While spaceborne observations indicate that this acceleration has levelled off recently (Rignot et al., 2019), they also support model projections suggesting modest changes...
in mass balance, i.e. the resulting net ice loss after accounting for all loss and gain processes, for the next decades to come (Bamber and Dawson, 2020). The dynamic ice loss is mainly responsible for the negative mass balance of Pine Island Glacier (PIG). The net input is commonly referred to as the surface mass balance (SMB), i.e. snowfall minus sublimation, meltwater runoff, and erosion/deposition of snow (Lenaerts et al., 2012; Medley et al., 2013). Various methods exist to measure the SMB on the ground (Eisen et al., 2008). The remoteness of WAIS makes such measurements logistically challenging, in particularly when extending these measurements to regional scales. Basin wide mass balance estimates strongly depend on the coverage and quality of SMB measurements. The study of Medley et al. (2014), in the following abbreviated as ME14, presents the first comprehensive survey of mean annual SMB between 1985 and 2009/10 from airborne radar based observations of the Thwaites and Pine Island glaciers. They demonstrated that such airborne radar observations provide a critical means to overcome the logistical challenges. However, these measurements rely on assumptions about the dielectric properties of snow, which include knowledge of the vertical snow density profile. In this sense, ground-truthing measurements remain an important tool for calibrating the radar soundings.

As part of the iSTAR Ice Sheet stability programme, a traverse across the Pine Island Glacier (PIG) was carried out in 2013/14 and repeated the year after. In the following, we name the first traverse T1 and the second T2. In total 22 sites were occupied during both traverses. Boreholes of at least 13 m depth were drilled at each site during traverse T1. Density–depth profiles were measured with a Neutron Probe (NP) device during both traverses (Morris et al., 2017) and supplementary analysis of firn cores was performed for 10 sites during traverse T2 to determine additional independent proxies related to the annual snow accumulation (Konrad et al., 2019).

The Alfred Wegener Institute (AWI) contributed to the iSTAR traverse T2 with radar soundings from the Airborne SAR / Interferometric Radar Altimeter System (ASIRAS) aboard the Polar 5 research plane. Previous ASIRAS missions have demonstrated its capability to trace annual snow accumulation layers of the upper firn column at regional scales over Greenland (Hawley et al., 2006; Overly et al., 2016). The PIG flight track connects all iSTAR sites so that internal annual snow accumulation layers can be traced to make regional scale SMB estimates. By comparison with earlier SMB measurements at PIG, the vertical profiling based on the ASIRAS soundings achieves a resolution that is one order of magnitude higher (Tab. 1), which helps to trace narrow internal snow accumulation layers. In addition, the ASIRAS flight track contains several self-intersections so that we can compare and adjust traced layers at those points.

In this study we first address local departures between SMB estimates from ASIRAS and NP measurement to evaluate the uncertainty of our regional scale ASIRAS SMB estimates. We then compare our results with those reported by ME14 and discuss differences between both data sets. Finally, we apply our new regional scale SMB estimates to different PIG mass balance inventories to evaluate their impact in light of the current stability of the study area.
Figure 1. ASIRAS-iSTAR survey projected on polar stereographic coordinates: black lines denote the ASIRAS flight track, numbered blue circles the iSTAR sites with shallow (∼13 m) neutron probe snow density measurements, and magenta-blue circles iSTAR sites with additional deep (∼50 m) firm core measurements of H$_2$O$_2$ during traverse T2. Surface flow speeds from Rignot et al. (2017) are overlayed by colour shadings on top of Landsat imagery (U.S., Geological Survey, 2007). Dotted lines denote basin outlines based on Fretwell et al. (2013) analysis (data accessed via the SCAR Antarctic Digital Database on 23 April 2019).

2 Data and methods

2.1 iSTAR traverse

The iSTAR traverse followed the PIG main trunk as well as its tributaries as shown in Fig. 1. A total flight track (black lines) of 2486 km was covered by the ASIRAS measurements between 1 and 3 December 2014. Following ME14, the basin outlines (dashed lines) include the Wedge zone between PIG and Thwaites. The main emphasis of the iSTAR campaign was on the fast flowing segments of PIG, thus we lack measurements from the southward interior. Earlier observations from ME14 suggest...
that the SMB decreases towards the interior so the contribution from this area to the total mass input will be less than that from the rest of the basin.

Additional SMB measurements were made with a ground penetrating radar during traverse T1 and published in Konrad et al. (2019). They selected the ∼1986 reflection layer, which approximately coincides with the observed main reflector by ME14, and traced the layer along sections of the 900 km traverse, amounting to a total of a 613 km distance covered by GPR observations. Due to the limited maximum sampling depth of the ASIRAS and NP measurements, the 1985/86 reflection layer used by ME14 and Konrad et al. (2019) is not contained in most of our data. To benefit from the ASIRAS coverage while simultaneously accounting for its limited depth range, we manually traced the continuous 2005 reflection layer over a distance of 2367 km to derive mean annual SMB estimates for the 2005–2014 period. Due to the reported consistency between the GPR and airborne SMB measurements in Konrad et al. (2019), we limit the comparison of our results to the basin wide estimates by ME14. In addition, we assume that the effect of strain history, which could affect our SMB estimates at the fast flowing sections of PIG, is negligible. Konrad et al. (2019) conclude that the total effect over the whole catchment is small, even though it can have a very significant effect at some sites. However, this effect is expected to be further reduced for the shallower reflection layer depths from the ASIRAS measurements.

### 2.2 Neutron Probe measurements

NP measurements of snow density were performed at all stations during both traverses as described in Morris et al. (2017). To evaluate the effect of densification, they repeated the density profiling in the same boreholes during traverse T2. Because the most recent accumulation is missing in these profiles, they drilled an additional borehole of less than 6 m depth and a nearby distance of about 1 m to capture it during traverse T2. The only exception is site 2, where the ground team decided to auger a completely new 14 m borehole for the density profiling due to poor data from the T1 hole. The deep firn cores (∼ 50 m) denoted in Fig. 1 were collected and analysed by the British Antarctic Survey. This analysis includes the annual variations with depth of the photochemical \( \text{H}_2\text{O}_2 \) tracer and density, which are phase shifted by about six months. According to Morris et al. (2017) the annual density variation is caused by alternating late austral summer/autumn low-density hoar layer with winter snow which has densified under the influence of warm summer temperatures. The different

| radar system       | ASIRAS          | CReSIS Accu-R  | pulseEKKO PRO |
|--------------------|-----------------|----------------|---------------|
| vertical range bin | 7.3 ± 0.3 cm    | 62 cm          | 100 cm        |
| along track bin    | 4.5 m           | O (10 m)       | 1.4 m         |
| maximum sampling depth | 30 m           | 300 m         | (90-120) m    |

Table 1. Approximate sample bin resolution and maximum depth of: SAR level_1b processed ASIRAS data with indicated vertical range bin standard deviation based on our two-way-travel time (TWT) to depth conversion, CreSIS Accumulation radar according to Medley et al. (2014), and pulseEKKO PRO GPR discussed in Konrad et al. (2019). For the GPR system we estimate the maximum sampling depths based on shared radargrams, which resolve the internal stratigraphy at PIG for TWTs up to (1000-1200) ns.
modulating processes of density and H$_2$O$_2$ allow for an independent determination of annual snow accumulation at the 10 deep core sites. Morris et al. (2017) applied an automatic annual layer identification routine to their snow density profiles and used the annual H$_2$O$_2$ peak depths as an additional guidance for the annual layer dating. Thus, the depth–age scales from both annual markers are consistent.

We use a single regional density–depth profile for the two-way-travel time (TWT) to depth conversion of the ASIRAS soundings, which we derive from the 43 NP profiles of traverse T2. First, we merge the $\sim$13 m and nearby $\sim$6 m density–depth profiles from the NP measurements during traverse T2 by linearly relaxing their overlapping segments. To reduce the effect of lateral noise convolution, we limit the relaxation length to the overlapping segments that correlate well with each other. Then we align the intercepting depth–age scales to create a consistent depth–age scale for each compiled profile. The resulting merged profiles are shown by the grey lines in Fig. 2. From these profiles, we then determine a smoothed regional mean profile, which is denoted by the black line. Finally we fit an exponential function to the regional mean profile (red dashed line), which we apply to the TWT to depth conversion. The blue dashed lines show the fitted standard deviation of the density as
Figure 3. ASIRAS radargram at iSTAR site 21 with surface reflection centred at the origin of the TWT scale: traced layer highlighted in magenta and depth–age scale at point of closest approach highlighted in cyan. The distance and trace numbers refer to the origin of the ASIRAS track segment 20156124 (see Tab. 2).

A function of depth. Following Medley et al. (2013), we consider the fitted standard deviation to be representative for the spatial uncertainty of the regional scale density–depth profile.

2.3 ASIRAS soundings

ASIRAS is a Ku-band radar altimeter which operates at a carrier frequency of 13.5 GHz and a bandwidth of 1 GHz (Mavromatidou et al., 2004). As in a previous campaign in Greenland (Overly et al., 2016), it was operated in Low Altitude Mode at PIG (i.e. less than 1500 m above ground). A Synthetic Aperture Radar (SAR)-processing of the collected data was performed, which yields the spatial resolution of the SAR level_1b data shown in Tab. 1. The associated cross track footprint is ~15 m.

We use the electromagnetic wave speed \( v = c/\sqrt{\epsilon'} \) to convert TWT to depth, where \( c \) is the vacuum speed of light and \( \epsilon' \) is the real part of the dielectric permittivity of the firn column. For the latter, we apply the commonly used empirical relation by Kovacs et al. (1995):

\[
\epsilon'_{\text{kov}} = (1 + 0.845 \rho_s)^2, \tag{1}
\]

where \( \rho_s = \rho/\rho_w \) is the specific gravity of snow at current depth with regard to the water density \( \rho_w = 1000 \text{ kg m}^{-3} \). An alternative model by Looyenga (1965) is

\[
\epsilon'_{\text{loo}} = \left( \frac{\rho}{\rho_{\text{ice}}} \left[ \sqrt{\epsilon'_{\text{ice}}} - 1 \right] + 1 \right)^3, \tag{2}
\]
with \( \varepsilon'_{\text{ice}} = 3.17 \) (Evans, 1965) and \( \rho_{\text{ice}} = 917 \text{ kg m}^{-3} \). Sinisalo et al. (2013), who consider a similar depth range to this study, conclude that the difference between wave speeds based on Eq. (1) and (2) has a negligible impact on their SMB estimates. This is also the case for our estimates (see Section 3). The maximum depth of the radargrams is \( \sim 30 \text{ m} \) based on the TWT to depth conversion from the fitted regional mean profile of density with depth and substituted Kovacs relation. The depth range of resolved internal stratigraphy varies along the flight track, but the layering remains visible for most of the upper 13 m depth covered by the NP measurements. Using the mean depth density profile, we determine the w.e. depth value for each waveform bin and calculate the mass per unit area between the selected reflection layer (magenta line in Fig. 3) and surface. We assume that the annual peaks in density lead to reflection layers because of the associated dielectric constant (Eisen et al., 2003). However, we also find that the ASIRAS soundings appear to resolve sub-annual layers at locations where the snow accumulation is high, e.g. around iSTAR site 21 as shown in Fig. 3. These layers may be created by intra-annual events, e.g. hoar events that form thin ice layers (Arcone et al., 2004), which generate a noticeable dielectric contrast with regard to the ASIRAS soundings. Hence, we rely on the independent measurements at the iSTAR sites to distinguish intra-annual from annual layers. Before the layer tracing, we apply an automatic-gain-control filter to all waveforms and limit their dynamic range to twice the standard deviation centred around the mean amplitude of each waveform. This improved the signal contrast of the radargram. Initially we tested a phase following algorithm of the Paradigm EPOS geophysical processing software to trace semi-automatically the selected reflection layer. However, this method became unstable for lower contrast and cases with close layer spacing. Furthermore, remaining SAR-processing artefacts were interfering with the phase following algorithm. Because of the complex nature of the observed stratigraphy, as has been also reported by Konrad et al. (2019), a manual layer tracing was used. Following Richardson et al. (1997), we attempted to bridge distorted or merged layer segments whenever distinct characteristics of a vertical layer sequence could be identified with confidence before and after the bridging. Different processes can lead to distortion of the reflection layers, e.g. processes changing deposition of the annual snow layers or excessive rolling angles of the airplane, while merging layers can result from low snow precipitation and ablative processes (e.g. wind-scouring) or a combination of both. We compared the traced layer at 34 intersections and 8 nearby flight track segments. In case of layer mismatches, the layer tracing was iteratively adjusted to find a consistent layer for the entire flight track. In this sense, the layer tracing is performed independently from the annual layer dating at each iSTAR site.

### 2.4 Measurement error estimation

We attempt to trace the 2005 reflection layer, which is covered by all NP depth-age scales. Following Morris et al. (2017), we define mass balance years between the density peaks in the NP profiles (nominally 1st of July). For instance, the mass balance year 2013 begins at the second annual density peak below the surface (nominally 1st July 2013) and ends at the first peak (1st July 2014). At \( N \) points along the closest approach of the ASIRAS track to each iSTAR site we estimated the depth of the reflection layer from its TWT, using the fitted density profile from that site. We also tested the use of the measured density profile for the TWT conversion instead, but we find that the impact on our results is negligible similar to Morris et al. (2017). The estimated depths do not necessarily coincide with the depth of a peak in the density profile i.e. the start of a mass balance year. We assign an estimated date to the mean value of \( N \) points by interpolation using the depth-age scale from the
| Track Number | iSTAR Site | Year       | $\Delta D$ [m] | Comment          |
|--------------|-----------|------------|----------------|------------------|
| 20156125     | 1         | 2003.9 ± 0.2 | 75            |                  |
| 20156125     | 2         | (2009.88 ± 0.03) | 8              | erroneous NP dating |
| 20156125     | 3         | 2003.39 ± 0.04 | 11            |                  |
| 20156125     | 4         | 2004.9 ± 0.1  | 86            |                  |
| 20156110     | 4         | 2004.5 ± 1    | 811           |                  |
| 20156125     | 5         | 2006.4 ± 0.1  | 119           |                  |
| 20156125     | 6         | 2005.1 ± 0.2  | 177           |                  |
| 20156106     | 6         | 2006.0 ± 0.6  | 2000          |                  |
| 20156125     | 7         | missing      | 226           | noise            |
| 20156115     | 7         | 2006.2 ± 0.4  | 304           |                  |
| 20156113     | 8         | 2004.8 ± 0.2  | 316           |                  |
| 20156109     | 8         | 2004.9 ± 0.2  | 470           |                  |
| 20156113     | 9         | 2003.1 ± 0.1  | 282           |                  |
| 20156113     | 10        | 2003.5 ± 0.4  | 9             |                  |
| 20156114     | 10        | 2004.9 ± 0.1  | 331           |                  |
| 20156114     | 11        | 2004.3 ± 0.2  | 121           |                  |
| 20156115     | 11        | 2004.3 ± 0.2  | 76            |                  |
| 20156115     | 12        | missing      | 490           | noise            |
| 20156115     | 13        | 2005.5 ± 0.2  | 127           |                  |
| 20156102     | 13        | 2005.7 ± 0.2  | 56            |                  |
| 20156109     | 13        | 2004.7 ± 0.1  | 377           |                  |
| 20156107     | 14        | 2003.5 ± 1    | 325           |                  |
| 20156109     | 14        | 2005.0 ± 0.6  | 350           |                  |
| 20156107     | 15        | 2005.16 ± 0.09 | 485         |                  |
| 20156120     | 15        | 2005.92 ± 0.2 | 190           |                  |
| 20156107     | 16        | 2004.19 ± 0.05 | 413        |                  |
| 20156103     | 16        | 2003.8 ± 0.1  | 314           |                  |
| 20156120     | 17        | missing      | 60            | noise            |
| 20156121     | 18        | 2004.33 ± 0.09 | 186        |                  |
| 20156120     | 18        | (2003.2 ± 0.2) | 297          | extrapolated     |
| 20156126     | 18        | (2002.0 ± 0.4) | 867          | extrapolated     |
| 20156121     | 19        | (1996.8 ± 0.3) | 230          | extrapolated     |
| 20156122     | 20        | 2005.12 ± 0.06 | 15           |                  |
| 20156122     | 21        | 2005.73 ± 0.02 | 115         |                  |
| 20156124     | 21        | 2005.71 ± 0.07 | 26           |                  |
| 20156124     | 22        | 2005.65 ± 0.06 | 21           |                  |

Table 2. Dating (Year) with associated uncertainty of closest ASIRAS reflection layer to $i^{th}$ iSTAR site. "Track number" refers to the ASIRAS flight track naming convention and "$\Delta D$" is the closest distance between the ASIRAS track and the iSTAR site. Years in brackets are discarded from the regional layer age estimation for the reasons given in the Comment column.
iSTAR site. These points are centred around the point of closest approach to the nearby iSTAR GPS location. We chose the associated interval length to be twice the distance $\Delta D$ between the iSTAR GPS location and point of closest approach of the flight track. Finally, we added six months to the estimated dates according to the mass balance year definition. This yields the estimated Year values for the reflection layer at each iSTAR site in Tab. 2. Following Konrad et al. (2019) we express the dating uncertainty as the standard deviation $\sigma_x$ in the $N$ measurements. In this sense, our error estimate is more conservative than the standard error of the mean. Furthermore, we assume that the uncertainty due to local variation in the stratigraphy is isotropic, which does not generally need to be true. However, according to Tab. 2 in most cases the overall impact of this effect is one order of magnitude smaller than the variability of dated years among all iSTAR sites. As indicated in Tab. 2, we excluded dating estimates around iSTAR sites 2 and 19. In both cases, our layer tracing revealed a large offset contrary to the neighbouring iSTAR sites. Possible reasons for these offsets could be systematic errors in the layer dating from the NP profiles, the variability

Figure 4. Spatial, temporal, digitization, and combined relative errors (left panels) and error partitioning (right panels). Grey background shades indicate the depth distribution of the traced 2005 reflection layer. (a,b) Based on error propagation according to Eq.(5). (c,d) Excluded error cancellation for the spatial error terms (see text).
of internal stratigraphy between the ASIRAS measurements and their closest approach to both iSTAR sites or systematic errors in the manual reflection layer tracing. The remaining exclusion of dating estimates at iSTAR sites 7, 12, and 18 is either due to high noise levels of the radargram or reflection layer depths exceeding the NP depth–age scales. Following Konrad et al. (2019) we estimate the reflection layer year by the mean of dating values at each site with an uncertainty of \( \Delta t = \sqrt{\delta t^2 + \delta \ell^2 + \ell_x^2} \), with the annual layer tracing uncertainty \( \delta t = \pm 1 \) years, the standard deviation of dating estimates \( \delta \ell \), and the propagated error \( \ell_x = 1/n \sqrt{\sum_i \sigma_x(i)^2} \) from the \( n \) lateral error estimates around each iSTAR site (\( i = \) site number), which we introduced in addition. The resulting reflection layer dating estimate is \( T = 2004.8 \pm 1.4 \), which corresponds to a layer age of \( a = 10.1 \pm 1.4 \).

The associated average surface accumulation rate \( \dot{b} \) in terms of w.e. depth per year is

\[
\dot{b} = \frac{1}{a \rho w} \sum_{i=1}^{m} \delta z_i \rho_i, \tag{3}
\]

where \( \delta z_i \) is the \( i^{th} \) depth increment of the radar waveform and \( \rho_i \) is the associated density. Substitution of the wave propagation speed for \( \delta z_i \) yields

\[
\dot{b} = \frac{1}{a \rho w} \sum_{i=1}^{m} \frac{c t_s}{\epsilon_i^j} \rho_i, \tag{4}
\]

where \( t_s = 0.37 \) ns is the ASIRAS vertical bin sampling time (i.e. 0.5 \( \times \) TWT per bin), and \( \epsilon_i^j \) refers to the permittivity value at the \( i^{th} \) bin. To avoid any confusion with previous summations, the final index \( m \) refers to the traced waveform bin at the reflection layer depth. It is evident from Eq. (4) that the spatial uncertainty of the density profile affects both the integration depth and incremental mass. Medley et al. (2013) and ME14 estimated the spatial uncertainty from the resulting SMB change by directly applying the standard deviation fits of their regional density profiles to the TWT to SMB conversion. Instead, we may propagate the error in Eq. (4), assuming that errors are uncorrelated and normally distributed. Based on the Kovacs relation according to Eq. (1) we account for the temporal, spatial, and digitization:

\[
\Delta \dot{b} = \frac{c t_s}{a \rho w} \left( \sum_{i=1}^{m} \left( \frac{\Delta \rho_i}{\epsilon_{kov,i}^j} \right)^2 \right) + \left( \frac{\Delta a}{a} \sum_{j=1}^{m} \left( \frac{\rho_j}{\epsilon_{kov,j}^j} \right)^2 \right) + \left( \frac{1}{3} \sum_{k=m-1}^{m+1} \left( \frac{\rho_k}{\epsilon_{kov,k}^j} \right)^2 \right), \tag{5}
\]

where \( \Delta a = \pm 1.4 \) years is the temporal uncertainty and \( \Delta \rho_i \) are the standard deviation intervals according to Fig. 2. Due to the small incremental density change of \(< 0.7\%\) along the entire profile, we approximate the digitization error by the mean SMB value of three consecutive bins centred at the final profile bin of the current integration depth. Figure 4 (a) displays the integrated individual measurement error components as well as the combined error according to Eq. (5) as a function of geometric depth. In addition, we include the error partitioning in (b). The grey background shades highlight the distribution of layer depths to visualise the relevant error range of our SMB estimates, which peaks around (5, 8, and 10) m. By comparison with Medley et al. (2013) and ME14, we find that our spatial error estimate based on Eq. (5) is reduced by about one order of magnitude while the standard deviation fits of their regional density profiles cover a similar range compared to ours. We may ignore the spatial error compensation in Eq. (5) by replacing the root-sum-of-squares (RSS) with absolute values: \( \sum_{i=1}^{m} \left( \Delta \rho_i / \epsilon_{kov,i}^j \right)^2 \rightarrow \left( \sum_{i=1}^{m} |\Delta \rho_i / \epsilon_{kov,i}^j| \right)^2 \). Hence, to comply with the studies above, we consider the more conservative spatial error propagation
Figure 5. Traced annual mean SMB between November 2004 and December 2014 from ASIRAS soundings with overlayed contour lines from a digital elevation model (see text) and landsat background imagery (U.S., Geological Survey, 2007). (a) High spatial resolution, iSTAR sites denoted by magenta/blue circles (deep cores) and white/blue circles (shallow cores). (b) Smoothed and downsampled SMB estimates as described in the text.

based on the sum of absolute values, but we keep the RSS of individual error components for the combined measurement error estimate as shown in Fig. 4 (c-d). Following these assumptions, we find that our measurement error estimate is still dominated by the temporal layer dating uncertainty for most of the traced layer depths, but the spatial error reaches a similar range as reported in Medley et al. (2013).

2.5 Kriging scheme

We focus on the regional scale variability of the SMB distribution on PIG. Figure 5 shows our high resolution (i.e. metre-scale) SMB estimates as well as smoothed SMB values with contour lines from a digital elevation model (DEM) by Helm et al. (2014). We use the same 25 km along-track smoothing window as ME14 and choose a sampling interval of half the smoothing...
Figure 6. Experimental semivariogram of log-transformed smoothed SMB observations (dots), Gaussian fit model (black solid line), sill and practical range parameter (dashed lines).

Figure 7. PP-plots between SMB observations and estimates based on OK and OLK interpolation methods for varying maximum estimate distances with regard to the closest measurement locations. The dashed 1:1 line indicates the complete PP-agreement between observations and estimates. Average PP-distances (see text) and SMB values are shown in the figure.

window length. We initially tested the same interpolation scheme as described in ME14 to estimate a regional scale SMB field for the PIG basin from our smoothed SMB points. This scheme is based on the ordinary kriging (OK) algorithm, a widely used
geostatistical interpolation technique (e.g. Isaaks and Srivastava, 1991). Instead of a direct OK interpolation of smoothed SMB observations, ME14 consider the residual SMB values with regard to an ordinary least squares linear regression model for the Thwaites-PIG basin area with northing, easting, and elevation as explanatory variables. This, in turn, yields a small degree of skewness < 0.5 of their new SMB distribution. However, we failed to reduce the skewness of residual SMB values effectively by the same method, which may be due to the different aerial coverage considered in our regression model. Examination of the DEM contour lines in Fig. 5 reveals that a simple relation between surface elevation and SMB is not evident, which may hint that the prevailing synoptic scale weather conditions at the Amundsen and Bellingshausen Sea sectors in combination with the precipitation shadowing effect of the Eights Coast mountain range (Fig. 1) require a more sophisticated model to capture the SMB at the PIG basin scale. An alternative approach, which is also mentioned in ME14, is a logarithmic transformation of the SMB observations prior to the OK interpolation:

$$\hat{B}(x_0) = \ln \left( \hat{b}(x_0) + C \right),$$

(6)

where $C$ is an arbitrary constant and $x_0$ represents the current interpolation location. After the OK interpolation of transformed SMB observations, the estimates must be transformed back into the original measurement scale. This backtransformation requires the addition of a nonbias term for each OK estimate to ensure that the expected value is equal to the sample mean and that the smoothing effect is adequately compensated (i.e. resulting estimates reproduce the sample histogram and sample mean [Yamamoto, 2007]). We implemented such ordinary logarithmic kriging (OLK) method in our analysis by adopting the 4-step post-processing algorithm proposed by Yamamoto (2007) for the estimation of nonbias terms. According to Yamamoto (2008), OLK does not necessarily require a log-normal sample distribution to produce improved estimates in terms of local accuracy. Furthermore, Yamamoto (2007) tested the impact of constant $C$ according to Eq. (6) and found that a data translation towards higher values yields an approximation from OLK to OK estimates, thus, eliminating the advantage of improved sample mean reproduction and local accuracy of OLK estimates. Indeed, we find that adding a negative constant $C$ to all SMB values, such that the lowest SMB value reaches 0.1 kg m$^{-2}$ yr$^{-1}$, yields an improved reproduction of the observation data characteristics. Figure 6 shows the experimental isotropic semivariogram of our log-transformed SMB observations from Fig. 5 (b) together with a Gaussian model fit with a practical range of $\sim$ 190 km, i.e. the range at which the spatial autocorrelation of sample points is vanishing. Following Yamamoto (2005, 2007), we investigate the reproduction of observational data characteristics by means of PP-plots (i.e. percentiles of cumulative distributions of observations and estimates against each other). Figure 7 shows the PP-plots for our OLK and OK interpolation constrained to a maximum estimation range with regard to the closest ASIRAS measurement locations of (100 and 190) km, and nearest neighbour locations. By comparison with Fig. 6, the (100 and 190) km distances (dashed lines) approximately correspond to the lag distances at which the semivariogram has reached half the sill and where it has levelled off respectively. In addition, the average distance of PP-points from the 1:1 line according to the definition in Yamamoto (2005) as well as the average SMB values for the OK and OLK estimates are shown in the legend. Both, the nearest neighbour OK and OLK average SMB estimates are close to the average SMB observation value of 474 kg m$^{-2}$ yr$^{-1}$. However, after extending the estimation range, it is evident from Fig. 7 that the best match exists between the observation and OLK estimation values. Hence, we limit our analysis to these values in the following.
Aside from the choice of the translational constant $C$ and semivariogram model, we choose the proposed method by Deutsch (1996) to correct for negative kriging weights (Yamamoto, 2000) and constrain all processing steps of the OLK estimation to the 16 nearest neighbours for each estimate according to the quadrant criterion. Depending on the considered neighbourhood the effect of smoothing as well as local stationarity of observation data is affected. As a guidance for our final setting, we aimed at generating an optimal PP-relation according to Fig. 7, but also considered potential artefacts, which may arise from the OLK procedure.

In addition to each OLK estimate, we calculate the associated interpolation error. While ME14 choose the kriging standard deviation as a measure of interpolation error, our error estimation is based on the introduced interpolation standard deviation $S_0$ by Yamamoto (2000) for two reasons. Firstly, as shown by the author, $S_0$ represents a more complete measure of local accuracy and has, therefore, been implemented in the post-processing algorithm in Yamamoto (2007). Secondly, for the OLK method we need a corresponding backtransformation of the interpolation error from the logarithmic to the measurement scale, which has already been investigated for $S_0$ in Yamamoto (2008). Thus, we adopted the proposed backtransformation of $S_0$ in this study.

Following ME14, we estimate the total error of each SMB estimate by the RSS of the measurement error and transform $S_0$ back. The measurement error is estimated by generating 500 realisations of OLK SMB estimates with add noise to the smoothed SMB observations, which follows a normal distribution with a mean of zero and standard deviation equal to the measurement error of the SMB observation at $x_0$.

For the basin wide SMB estimation we have to keep in mind that our OLK estimation is limited in terms of the practical range according to Fig. 6. By comparison with the flight track shown in Fig. 1, even when considering the practical range as a maximum threshold for the spatial SMB estimation, we do not cover the entire PIG basin. Hence, for the calculation of total mass input for the PIG basin, we replace our SMB estimates with modelled SMB from the RACMO2.3p2 (van Wessem et al., 2018) regional climate model (in the following abbreviated as RACMO) at distances where the spatial autocorrelation of measurements is low.

3 Results

3.1 Regional scale SMB distribution

Based on the adopted OLK interpolation scheme, we produced the mean annual SMB map for the PIG basin from the ASIRAS observations in Fig. 8 (a). SMB observations and estimates are colour coded with the same scale. Each estimate covers a pixel size of $\sim 5$ by 5 km$^2$ and refers to the averaging period between November 2004 and December 2014. The two surrounding dashed lines indicate the (100 and 190) km maximum distances from the ASIRAS measurement point cloud as discussed earlier. The red triangle denotes an artificial interpolation cluster of 8 pixels with SMB values greater than 2000 kg m$^{-2}$ yr$^{-1}$, which we will discuss in Sec. 4.3.

By comparison with ME14 and the extracted RACMO estimates in Fig. 8 (b), the regional SMB distribution at PIG reproduces the same main characteristics, i.e. increasing SMB rates towards the Amundsen Sea coastline, decreasing SMB rates towards...
Figure 8. (a) ASIRAS annual mean SMB OLK estimates between November 2004 and December 2014. Measurement points colour coded with the same SMB scale. Red triangle denotes the position of an interpolation artefact (see Sec. 4.3). (b) RACMO annual mean SMB estimates for the same averaging period as in (a). ASIRAS measurement locations denoted by dots. (c) Relative SMB change from RACMO to ASIRAS, i.e. 
\[
\frac{(a) - (b)}{b} \times 100.
\]
(d) Hybrid SMB map with (a) linearly relaxing into (b) between dashed lines (a,c,d), which denote the (100 and 190) km maximum distance to the ASIRAS measurements. Background imagery taken from U.S., Geological Survey (2007) for all panels.

the inland, and a region of low SMB in response to the shadowing effect from the Eights Coast mountain range. The RACMO estimates indicate a sharp boundary of this effect at the ice divide at the northern tip of PIG. This boundary is also captured by the ASIRAS observations at its northernmost flight track near iSTAR site 10, which is best seen in the high resolution observations according to Fig. 5 (a).

Fig. 8 (c) shows the relative difference between RACMO and ASIRAS estimates as defined in the caption. It is evident that the shadowing effect is more pronounced in the ASIRAS observations, while the southward SMB gradient is more pronounced in....
the RACMO estimates. Despite the missing ASIRAS observation towards the southern interior, the latter is already apparent inside the 100 km maximum distance margin (inner dashed line). By comparison with the SMB distribution according to ME14, we find similar departures to our results with respect to the northern slopes of PIG and its southward interior.

The ME14 results indicate an elevation dependent drift between observed and simulated SMB values, which we also find between our results and RACMO as shown in Fig. 9. Here, we first calculated the average SMB for each RACMO pixel from collocated ASIRAS estimates. Then we subtracted the RACMO SMB estimates from the ASIRAS averages and plot the resulting difference against their associated elevation, which we determined from the DEM model in Fig. 5. Due to the limited practical range of the ASIRAS observations, we only consider ASIRAS–RACMO estimates within the 100 km maximum range limit to the observations inside the PIG outlines. Like ME14, we find for this version of RACMO that simulated SMB estimates appear to be lower than the ASIRAS estimates at higher elevation levels and vice versa. Despite the scatter between elevation and SMB difference, a Kendall’s rank correlation test with \( \alpha = 0.05 \) significance level suggests that both are significantly correlated with each other.

### 3.2 Total mass input

For the calculation of spatially integrated SMB, denoted by \( \Sigma_+ \) in the following, we produced the hybrid SMB map in Fig. 8 (d), where the ASIRAS SMB estimates are linearly relaxed into the RACMO SMB estimates between the 100 km and 190 km range limits to account for the spatial autocorrelation in terms of Fig.6. Table 3 summarizes \( \Sigma_+ \) and further statistical SMB characteristics for different data sets and basin definitions according to Fig. 10 (d). Here, we replaced the interpolation artefact highlighted in Fig. 8 (a) with averaged values from neighbouring pixels. \( \Sigma_+ \) uncertainty estimates refer to the RSS of the
interpolation and measurement error-grids (Fig. 10 d) in accordance with ME14. Because RACMO is missing an error-grid, we consider the total combined error for the entire PIG basin as a conservative error estimate. In this sense, we are augmenting the missing model error estimation. To quantify the relative contribution of ASIRAS to the hybrid SMB estimates, OLK area and OLK $\Sigma_+$ denote the relative contribution in terms of covered land area and integrated SMB respectively. For comparison with our hybrid based estimates, we include results from RACMO and ME14, which we converted from w.e. depth to SI units. Because of the different averaging periods between this study and ME14, we add RACMO estimates in brackets, which we extracted based on the same averaging period as the ME14 results.

Table 3. Spatially integrated SMB ($\Sigma_+$), mean ($\mu_+$), standard deviation ($\sigma_+$), minimum, and maximum SMB based on different data sets and basin definitions according to Fig. 10 (d). For the hybrid SMB estimates of this study, the areal contribution as well as the contribution of spatially integrated SMB from the ASIRAS estimates is denoted by OLK Area and OLK $\Sigma_+$ respectability. In addition to the November 2004 to December 2014 averaging period of the hybrid estimates, RACMO estimates in brackets refer to the July 1985 to January 2010 averaging period in accordance with the results from Medley et al. (2014).

285 Pine Island and Wedge Zone

The Pine Island $\Sigma_+$ values are in agreement between all data sets within the estimated error margins. This is different for the Wedge outlines, where the RACMO $\Sigma_+$ estimates are between 35-40% lower compared to the estimates of this study and ME14. Increasing the averaging time of RACMO estimates to the 1985–2009/10 period of the ME14 results yields an increase of $\Sigma_+$ by 2% for the Pine Island and 8% for the Wedge outlines. In comparison with the ASIRAS and ME14 estimates, the simulated total mass input for the Wedge outlines remains outside the error margins of both observations. Considering the further SMB properties according to Tab. 3, the hybrid SMB estimates show the largest variability and range for the Pine
Island outlines. This is different for the Wedge outlines, where the hybrid SMB estimates appear to be larger for most of the basin area.

**Additional basin definitions**

Tab. 3 includes results based on two additional basin definitions for PIG. Figure 10 (d) shows a composite plot of all basin definitions used here. The surface areas range between 176.5, 178.6, and $208.8 \times 10^3$ km$^2$ for the PIG basin (including Wedge) according to the definitions of Mouginot et al. (2017), Fretwell et al. (2013), and Zwally et al. (2012). Alongside with extended surface cover according to Zwally et al. (2012), $\Sigma_e$ estimates increase between 16 to 19 % with regard to the other basin definitions. The impact on the mean annual SMB is limited to 10 kg m$^{-2}$yr$^{-1}$ ($\sim 2\%$) between all catchment definitions.

4 Discussion

We discuss first the pronounced differences between annual layer dating from ASIRAS reflection and neutron probe density profiles at some sites and then secondly the systematic differences in SMB distribution between our results and those of ME14 and RACMO.

4.1 SMB departures

Key to the evaluation of our selected internal reflection layer is its isochronic nature, which we assume based on matched depth–age relations from the iSTAR ground truthing measurements. One may argue that these measurements can be subject to local noise in the density profile, which would challenge any comparison with nearby radar observations. For instance, Laepple et al. (2016) observed dominating stratigraphic noise at single pit density profiles near Kohnen station (East Antarctic plateau, Dronning Maud Land). Stacking of multiple profiles is one possibility to filter out noise. While this is not possible for the single iSTAR sites, the estimated dating uncertainty of $\pm 1.4$ years according to this study suggests that iSTAR ground truthing measurements at PIG are less prone to stratigraphic noise, which is most likely to be related to the higher SMB compared to $\sim 70$ kg m$^{-2}$yr$^{-1}$ nearby Kohnen station (Laepple et al., 2016). However, on a few occasions we identified larger departures in the annual layer dating, as it is the case for iSTAR site 2 (Tab. 2). While the layer tracing appears to be in agreement between site 1 and 3, the annual layer dating at site 2 would suggest an SMB of $\sim 290$ kg m$^{-2}$yr$^{-1}$ at the traced layer cross section rather than $\sim 150$ kg m$^{-2}$yr$^{-1}$ based on the 2004.8 layer dating of this study. Accordingly, local SMB results would increase by $\sim 100\%$, if we used the uncorrected depth–age scale at site 2, which most likely indicates a systematic error in the measurement scale. This is further corroborated by the measured SMB of 140 kg m$^{-2}$yr$^{-1}$ at site 2 for the most recent 2014 layer, but also measured density and strainrate profiles suggest a mean annual SMB of 200 kg m$^{-2}$yr$^{-1}$ based on theoretical grounds, which both are in a better agreement with the collocated ASIRAS based results. In this sense, the ASIRAS results allow us to be more confident of the site 2 strainrate measurements and therefore add to the densification analysis of Morris et al. (2017). The local SMB estimates near site 2 from ME14 and RACMO are within the 200 to 300 kg m$^{-2}$yr$^{-1}$ range, but lack the local precision of ASIRAS measurements and therefore could not explain the measured density and strainrate profiles.
Table 4. Updated mass balance estimates $\Sigma_+^-$ for different studies based on hybrid $\Sigma_+^+$ estimates of this study. Indicated periods refer to the considered ice loss processes. Net gain from $\Sigma_+^+$ is assumed to be stationary during the indicated periods.

| Study                | Period      | Basin Definition                  | $\Sigma_+^-$ (Gt yr$^{-1}$) | Updated $\Sigma_+^-$ (Gt yr$^{-1}$) |
|----------------------|-------------|-----------------------------------|----------------------------|-----------------------------------|
| Medley et al. (2014) | 2005 - 2012 | Pine Island (Medley et al., 2014) | $-41 \pm 7$                | $-39 \pm 18$                     |
| Medley et al. (2014) | 2005 - 2010 | Wedge (Medley et al., 2014)       | $-0.6 \pm 2.5$             | $-0.7 \pm 3.1$                   |
| Gardner et al. (2018)| 2008 - 2015 | Zwally et al. (2012)              | $-49 \pm 19$               | $-40 \pm 28$                     |
| Rignot et al. (2019)| 2005 - 2014 | Mouginot et al. (2017)            | $-51 \pm 7$                | $-46 \pm 20$                     |

4.2 Impact on recent mass balance estimates

Despite the local differences in the SMB distribution, the difference between the $\Sigma_+^+$ estimates for the PIG catchment (including Wedge) between this study and ME14 is small, i.e. our $\Sigma_+^+$ is greater by $1.7$ Gt yr$^{-1}$, which corresponds to $2\%$ of the ME14 value. This indicates that the local differences in the SMB estimates between both studies cancel out. If we take into account that the temporal averaging time used by ME14 is about a factor of $2.7$ larger than that used in this study, this indicates that the observed total mass input for the PIG basin is not subject to any secular trend throughout the observational period, a finding which is further corroborated by the iSTAR measurements and RACMO simulations. This provides additional evidence to Medley et al. (2013) that the recent temporal evolution of the PIG mass balance is primarily driven by dynamic ice
Table 5. Mean $\mu_{err}$, standard deviation $\sigma_{err}$, minimum, and maximum gridded SMB errors. Mean and $\sigma$ values in brackets refer to the catchment area within the 190 km practical range limit and are weighted in terms of the ASIRAS partitioning of the final hybrid SMB map in Fig. 8 (d).

| %     | Pine Island | Wedge |
|-------|-------------|-------|
|       | $\mu_{err}$ | $\sigma_{err}$ | min | max | $\mu_{err}$ | $\sigma_{err}$ | min | max |
| Measurement Error | 9.9 (7.2) | 5.2 (1.8) | 3.7 | 40.5 | 7.2 | 0.9 | 5.8 | 9.5 |
| $S_0$ | 32.2 (25.5) | 32.7 (34.9) | 0.5 | 481.8 | 12.5 | 11.3 | 2.1 | 57.3 |
| Combined Error | 34.4 (27.4) | 32.4 (34.2) | 0.6 | 481.9 | 15.2 | 10.2 | 7.3 | 57.8 |
| Medley et al. (2014) | 10.4 | 6.1 | 2.6 | 30.0 | 6.3 | 1.8 | 2.9 | 10.8 |

With regard to existing mass balance estimates for PIG, we have to take into account that the definition of basin can differ significantly as illustrated in Fig. 10 (d). To evaluate the impact of our hybrid SMB estimates on recent mass balance inventories, we extracted results from the literature in Tab. 4 and added updated mass balance estimates $\Sigma^+_+ \Sigma$ by replacing the $\Sigma^+_+$ estimates from the literature with the $\Sigma^+_+$ estimates of this study. For the periods shown, we assume that the SMB remains stationary for the mass balance calculation. In addition, we linearly interpolated the estimated ice discharge measurements in ME14 for the missing periods before 2007. Furthermore, we assume that the unspecified basin definitions in ME14 are in close agreement with the basin definitions based on Fretwell et al. (2013).

The small difference between the $\Sigma^+_+$ estimates of this study and ME14 directly translates into the $\Sigma^-_+$ mass balance estimates. The largest impact of our results is on the $\Sigma^-_+$ estimate by Gardner et al. (2018). After replacing their $\Sigma^+_+$ estimate from RACMO2.3 simulations with our hybrid $\Sigma^+_+$ estimate, the mass balance increases by $\sim 9 \text{ Gt yr}^{-1}$.

4.3 SMB uncertainty

While the agreement in $\Sigma^+_+$ estimates between this study and ME14 supports the hypothesis that the regional SMB of PIG is stationary at decadal scales, our uncertainty estimates are much larger. The temporal error according to Fig. 4, which is $\sim 5\%$ larger than Medley et al. (2013) and ME14, cannot fully explain the difference between both uncertainty estimates. We also do not expect any major differences with regard to the spatial uncertainty of the density profiles. According to the error-grid statistics in Tab. 5, we identify the backtransformed interpolation standard deviation $S_0$ from the OLK scheme as the dominating error source of our results, while the combined error in ME14 is slightly above our measurement error. The dominating $S_0$ uncertainty is also evident in Fig. 10 (a,b,c), where the spatial features of the combined error-grid are predominately determined by the $S_0$ grid. We find that the low accumulation zone at the northern slopes of PIG, which is next to the main trunk between iSTAR site 1 to 6, shows combined $S_0$ patches that considerably exceed 100%. In contrast, combined error estimates in ME14 do not exceed 20% at the same location.

Initial tests on our OLK setting revealed that the choice of the negative kriging weights correction method has a noticeable
Figure 10. (a) Interpolation standard deviation $S_0$ of Fig. 8(a). (f) Measurement error of Fig. 8(a). (c) Root sum square of (a,b). (d) Varying PIG basin definitions according to Zwally et al. (2012), Fretwell et al. (2013), and Mouginot et al. (2017). Surface flow speeds adopted from Fig. 1. Background imagery taken from U.S., Geological Survey (2007) for all panels.

impact on the uncertainty estimates, a finding, which according to our knowledge, has not been reported before. However, our
applied method by Deutsch (1996) already yields the minimum uncertainty estimates for our results, whereas the additional
methods cited in Yamamoto (2000) yields an additional uncertainty increase between 20% (Froidevaux, 1993) and 50% (Jour-
nel and Rao, 1996).

Additional tests, where we used the kriging standard deviation based on non-transformed OK estimates, did not improve our
interpolation uncertainty. Therefore the different choice of the interpolation uncertainty measure is not the source of the larger
uncertainty range of this study. We hypothesize that despite the homoscedastic (i.e. data value independent) nature of the krig
standard deviation, the reduction of data variance after subtracting the regression surface according to ME14 is most likely the
cause of their significantly lower uncertainty estimates.

In addition to the larger uncertainty range of this study, we note that the choice between cell-by-cell summation and RSS of grid errors has a quite substantial impact on the \( \Sigma_+ \) uncertainty estimates. If we make the optimistic assumption that gridded errors are independent and choose the calculation of RSS instead, \( \Sigma_+ \) uncertainty estimates would reduce to \( \pm 0.5 \text{ Gt yr}^{-1} \) (i.e. \( \sim 97\% \) less) for the combined Pine Island and Wedge basin.

4.4 Systematic retrieval impacts

In addition to the uncertainty assessment in section 4.3, we evaluated the impact of artificial cluster removal, the choice of the permittivity model, and the non-transformed OK scheme.

Artificial cluster removal

Inspection of the artificial cluster highlighted in Fig. 5 revealed that it is centred around the location with the lowest observed SMB and is essentially generated by the local nonbias terms of the OLK procedure. Owing to its steep contrast with the surroundings, it appears to be plausible to replace this cluster by averaged values of its nearest neighbours. However, due to the limited extend of this cluster, its additional contribution to the \( \Sigma_+ \) estimates would be less than 0.8%. Similarly, the impact on the PP-plot is negligible. Increasing the translational constant \( C \) helps removing this cluster, but at the cost of statistical agreement between observations and estimates.

Looyenga based results

Defining \( \epsilon' \) by Eq. (2) instead of Eq. (1) yields a minor reduction of \( \Sigma_+ \) for the PIG catchment of 0.6%, which we expect from Sinisalo et al. (2013). However, despite the minor impact of the alternative definition for \( \epsilon' \), we noticed an additional small impact on the layer dating, which shifted our estimated layer formation from November to September 2004. Thus, we had to adjust the time range in the RACMO SMB extraction for the calculation of hybrid SMB estimates. While the choice of the \( \epsilon' \) model only has a minor impact on our total mass input estimates, it is worth noting that the effect on our annual layer dating is detectable.

Non-transformed kriging results

If we choose the OK procedure instead, \( \Sigma_+ \) increases by 4% for the Pine Island and 12% for the Wedge zone outlines, which would further increase the offset between this study and ME14. However, inspection of the SMB distribution (not shown) indicates that estimates tend to overshoot near the coastline of the Amundsen Sea, which becomes particularly evident for the Wedge outlines. Hence, the OK procedure appears to be more sensitive to the limited observational constraints near the Wedge outlines. In addition, \( S_0 \) based uncertainty estimates increase by 27% and 88% for the Pine Island and Wedge outlines, which highlights the improved performance of the OLK procedure.
5 Conclusions

Our analysis provides updated mean annual SMB estimates for the PIG basin and 2005–2014 averaging period based on a comprehensive airborne radar and ground truthing survey and complementary model simulations. Based on these estimates, we calculated a total mass input of \( 79.9 \pm 19.2 \text{ Gt yr}^{-1} \) for the PIG basin area. In comparison with earlier estimates from airborne radar observations, which consider the 1985–2009 averaging period, our result shows 2% greater total mass input. Given the uncertainty in both values, there is no significant difference between them. Hence, no distinct secular trend is visible between both averaging periods. We conclude that our results provide further evidence that the recent total mass input can be considered stationary at decadal scales. This implies that the increased dynamic ice loss over past decades remains the driving force in the recent mass balance evolution of PIG. However, departures between both observations at the northern slopes and southward interior of PIG may indicate temporal changes in the local SMB distribution, which cancel out for the estimates on total mass input. Furthermore, our radar based observations can resolve a discrepancy between strainrate and SMB measurement at iSTAR site 2, which highlights the benefit of such complementary SMB measurements for future missions.

Despite the minor changes in total mass input between both studies, the more than twofold uncertainty range of our results remains striking. Neither can the applied model for the wave propagation speed of radar soundings nor the uncertainty related to the regional density profile explain the larger uncertainty of this study. The same also applies for the reduced temporal averaging time. A comprehensive evaluation of our uncertainty estimation revealed that assumptions on the geostatistical interpolation error as well as grid-error dependences can have a substantial impact on the uncertainty estimation. In terms of the error partitioning, our interpolation error is the dominating source of combined grid-errors. Moreover, varying basin definitions have an impact on our total mass input estimate by up to 19%. This highlights the importance of a thorough documentation of uncertainty estimates and basin definitions to improve future intercomparisons between different SMB and mass balance inventories.

Data availability. data submission to PANGAEA (in preparation)

Appendix A

A1

Author contributions. S. Kowalewski conceived of the presented idea, designed the computational framework, adapted and tested the geostatistical kriging methods, accomplished the reflection layer tracing in large parts, reprocessed the Neutron Probe density profiles for the data calibration, wrote the manuscript with input from all authors, V. Helm performed the SAR level_1b ASIRAS data processing, provided the digital elevation model, and established access to the RACMO2.3p2 data, E. Morris delivered the Neutron Probe density profiles, and O. Eisen contributed to layer analysis and interpretation. All authors discussed the results and contributed to revising the manuscript.
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