Physical activity and genetic predisposition to obesity in a multiethnic longitudinal study

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Physical activity (PA) has been shown to reduce the impact of FTO variation and obesity genetic risk scores (GRS) on BMI. We examined this interaction using a quantitative measure of PA and two adiposity indexes in a longitudinal multi-ethnic study. We analyzed the impact of PA on the association between 14 obesity predisposing variants (analyzed independently and as a GRS) and baseline/follow-up obesity measures in the multi-ethnic prospective cohort EpiDREAM (17423 participants from six ethnic groups). PA was analyzed using basic (low-moderate-high) and quantitative measures (metabolic equivalents (METS)), while BMI and the body adiposity index (BAI) were used to measure obesity. Increased PA was associated with decreased BMI/BAI at baseline/follow-up. FTO rs1421085, CDKAL1 rs2206734, TNNI3K rs1514176, GIPR rs11671664 and the GRS were associated with obesity measures at baseline and/or follow-up. Risk alleles of three SNPs displayed nominal associations with increased (NTRK2 rs1211166, BDNF rs1401635) or decreased (NPC1 rs1805081) basic PA score independently of BMI/BAI. Both basic and quantitative PA measures attenuated the association between FTO rs1421085 risk allele and BMI/BAI at baseline and follow-up. Our results show that physical activity can blunt the genetic effect of FTO rs1421085 on adiposity by 36–75% in a longitudinal multi-ethnic cohort.

Obesity has become a global epidemic1 and is a known risk factor for a number of adverse health outcomes, including psychological disturbance, osteoarthritis, type 2 diabetes, hypertension, cardiovascular disease, cancer and ultimately 8–13 years shorter life expectancy in its more severe forms2,3. Since the rising prevalence of obesity has been attributed primarily to changes in environmental exposures, such as excessive energy intake, sedentary lifestyle and sleep debt, among others4, recent research has focused on preventative strategies to control the obesity epidemic5.

Despite the global impact of these environmental changes, obesity appears to manifest preferentially in genetically predisposed individuals, and a high level of inter-individual variation has been observed among exposed populations5. Current evidence has shown that heritability estimates for obesity-related traits can be modulated by lifestyle factors such as physical activity (PA)7. Significant gene-environment interactions (GEI) between FTO intron 1 variation and PA have been consistently found in 16 independent cross-sectional and intervention studies performed with children and adult populations of European, East Asian and African ancestry8–10. A recent meta-analysis in 111,421 subjects of European ancestry confirmed a significant PA x genetic risk score (GRS) interaction using 12 obesity predisposing SNPs and showed that this interaction was more apparent in subjects living in North America11.

Although these data provide convincing evidence of an interaction between genetic predisposition to obesity and PA, we identified several important limitations of the existing literature. First, almost all current gene x PA interaction studies in the obesity field have been conducted using cross sectional designs, although there are exceptions12. Strengths of the cohort design include the ability to study temporal relationships, which strengthens
the confidence in statistical associations\textsuperscript{13,14}. Second, most GEI analyses of PA have used imprecise self-report measurements of PA due to concerns regarding cost, feasibility and participant burden\textsuperscript{15} and only one study has examined the impact of a quantitative metabolic equivalent (MET) score\textsuperscript{16}. Third, nearly all GEI studies of obesity have used BMI as the outcome measure of obesity, although the validity of the BMI can be compromised by different body compositions (e.g. lean vs. fat mass)\textsuperscript{17}. Fourth, GEI studies have mainly been performed in European populations, but the transferability of the conclusions of these studies to other ethnic backgrounds remains unclear\textsuperscript{18,19}. Although a previous meta-analysis analyzed multiple ethnic groups, the sample was over 95% European\textsuperscript{18}. This prompted us to assess (1) the association between PA behavior and obesity and (2) the interaction between obesity predisposing variants (analyzed independently and as a GRS) and PA on obesity-related traits in the multi-ethnic longitudinal cohort EpiDREAM. We used a quantitative measure of energy expenditure (Metabolic Equivalent Score, MET score)\textsuperscript{20} and the recently validated body adiposity index (BAI)\textsuperscript{21}. The BAI may provide new insight into this interaction due to its strong correlation ($r = 0.85$) with body adiposity (as measured by dual x-ray absorptiometry)\textsuperscript{21}.

**Materials and Methods**

**Study Participants.** The data for this investigation were collected through a prospective cohort of participants at risk for type 2 diabetes (T2D), which has been described in detail previously\textsuperscript{22,23}. Briefly, EpiDREAM enrolled a total of 24 872 individuals recruited from 21 countries who were screened for eligibility to enter the DREAM clinical trial\textsuperscript{22}. Individuals who were identified as at risk for type 2 diabetes based on family history, ethnicity and abdominal adiposity were screened using a 75-g oral glucose tolerance test (OGTT). A subset of 1960 DREAM participants received rosiglitazone over follow-up. All participants were between the ages of 18–85 years and were screened between July 2001 and August 2003. We focused on 17 423 subjects from six ethnic groups (South Asian, East Asian, European, African, Latin American, Native North American) having both phenotypic and 50 K gene-centric array information in the EpiDREAM study (Supplementary Figure 1). Self-reported ethnicity has been verified in the 17 423 individuals using the eigensof software (http://genepath.med.harvard.edu/~reich/Software.htm) and 40 individuals were reclassified. The first 10 components from this principal components analysis were retained to adjust for population stratification (Supplementary Figure 2).

Of the 17 423 participants at baseline, our follow-up analyses included 9 228 participants with complete genotype and phenotype data. Interim and final visits with participants occurred 12–24 months and 36–48 months (median follow-up 3.3 years) after their screening visit, respectively. Participants who were not able to complete in-person clinic visits were contacted through telephone or mail\textsuperscript{23}. The EpiDREAM study has been approved by local ethics committees and the study methodology was carried out in accordance with the approved guidelines. Informed consent was obtained from each subject before participating in the study, in accordance with the Declaration of Helsinki.

**Genotyping.** Buffy coats for DNA extraction have been collected in 19 197 participants of the EpiDREAM study (Supplementary Figure 1). DNA has been extracted by the Gentra System. Genotyping was performed using the Illumina CVD bead chip microarray ITMAT Broad Care (IBC) array\textsuperscript{24}. Genotyping was performed at the McGill University and Genome Quebec Innovation Centre using the Illumina Bead Studio genotyping module, version 3.2. We established a list of SNPs that reached genome-wide significance (P < 5 × 10\textsuperscript{-8}) with BMI or binary obesity status in populations of European ancestry. We used three different strategies to optimize the SNP selection procedure using a key word search (e.g. BMI) on i) the National Human Genome Research Institute (NHGRI) GWAS Catalog (www.genome.gov/gwastudies/) ii) the HuGE Navigator GWAS Integrator (www.hugenavigator.net/HuGENavigator/gWASHitStartPage.do) iii) the PubMed database (www.ncbi.nlm.nih.gov/pubmed). Using this strategy, we ended-up with a list of 72 independent SNPs associated with obesity-related traits. From this list, 14 SNPs were available on versions 1 and 2 of the IBC 50 K SNP array (Supplementary Table 1). The SNPs selected showed no significant (P < 10\textsuperscript{-6}) deviation from Hardy-Weinberg Equilibrium (HWE) in the six ethnic groups. The call rate for each of the 14 SNPs was between 99.8–100% (Supplementary Table 1).

**Phenotyping.** In addition to the OGTT, participants completed a questionnaire that included demographic data, medical history and PA behaviors at baseline and follow-up. Anthropometric measurements including height, weight, waist and hip circumference were performed using a standardized protocol\textsuperscript{25}. Height (m), weight and hip circumference (HC) (cm) were measured by trained medical staff. Standing height was measured to the nearest 0.1 cm and weight was measured to the nearest 0.1 kg in light clothing. Hip circumference was measured in duplicate at the level of the greater trochanters using a non-flexible tape measure with an attached spring balance with a mass of 750 g. Averages of the two measures were used in all analyses. Body mass index (BMI) was calculated as weight in kilograms divided by height in meters (m) squared. We used Body Adiposity Index (BAI), which estimates body adiposity percentage directly based on height and hip circumference\textsuperscript{21}. Specifically, $\text{BAI} = ([\text{hip circumference}] / [(\text{height})^{1.5}] - 18)$, with hip circumference and height being expressed in centimeters and meters, respectively\textsuperscript{25}.

The 2003 ADA criteria were used to classify participants as having normal glucose tolerance (NGT), impaired fasting glucose (IFG), impaired glucose tolerance (IGT), or T2D at baseline, as confirmed by an oral glucose tolerance test. Norgomycemia was defined as a fasting plasma glucose < 5.6 mmol/L, IFG was defined as a fasting plasma glucose of 5.6 to 6.9 mmol/L, IGT was defined as a fasting plasma glucose below 7.0 mmol/L and a 2-h glucose between 7.8 and 11.0 mmol/L, and diabetes was defined if either the fasting plasma glucose was ≥ 7.0 mmol/L or the 2-h glucose was ≥ 11.1 mmol/L\textsuperscript{26}. IFG and IGT were collapsed into one category and these three groups (norgomycemia, IFG/IGT, diabetes) comprised the glycemic status variable.
PA measures were based on self-reported time of participation in 41 different physical activities (Supplementary Table 2). A quantitative measure of energy expenditure (Metabolic Equivalent score, MET score) was derived using this information and the updated compendium of physical activities. A MET score was assigned to each activity based on the energy cost of that activity. The energy expended on each leisure time and work-related activity was estimated by multiplying the hours/week of participation by the corresponding MET value. These measures were summed across all activities to provide an overall estimate of energy expenditure represented as MET minutes per week. This quantitative measure was compared to a brief, self-reported, categorical measure based on PA participation during leisure time and at work. Participants rated their PA level at work and during leisure on a scale of one to three to form the basic PA score (1 = sedentary, 2 = moderately active, 3 = active).

Statistical Analyses. Statistical analyses were performed using SPSS (version 20, New York, USA, IBM Corporation) and the power of our study was calculated using QUANTO (version 1.2.4, University of Southern California, Los Angeles, CA, USA). Single SNP analyses were performed under the additive model, and the obesity risk alleles previously identified for each of the 14 SNPs in literature were used as the risk allele. General linear models (GLM) were used to examine (1) the association between energy expenditure (METs)/basic PA score at baseline and BMI/BAI at baseline and follow-up (2) the association between 14 obesity predisposing variant and environmental exposure can bias the estimates of the main genetic effect and the gene-environment interaction. Previous examples of these effects have been observed in cancer research, where variants on 15q25 have been linked to both smoking behaviors and lung cancer. The genetic predisposition score was calculated by summing the alleles of the 14 obesity predisposing SNPs so that the score could range from 0 to 28. Since weighting has been shown to have no major impact on the effect of a GRS, an unweighted GRS was used. We performed imputations for missing genotypic values as previously described using the mean number of predisposing obesity alleles in successfully genotyped individuals. This procedure was performed separately in each ethnic group. Individuals with more than one out of 14 missing genotype were not included in the genetic risk score calculation. Although multiple SNPs included in the gene score are within the same gene, our linkage disequilibrium analyses of this cohort indicated that it is appropriate to include all 14 SNPs in the genetic risk score (adjusted for sex, age, ethnicity, glycemic status and BMI). This test was performed since a correlation between the genetic variant and environmental exposure can bias the estimates of the main genetic effect and the gene-environment interaction. Previous examples of these effects have been observed in cancer research, where variants on 15q25 have been linked to both smoking behaviors and lung cancer. The genetic predisposition score was calculated by summing the alleles of the 14 obesity predisposing SNPs so that the score could range from 0 to 28. Since weighting has been shown to have no major impact on the effect of a GRS, an unweighted GRS was used. We performed imputations for missing genotypic values as previously described using the mean number of predisposing obesity alleles in successfully genotyped individuals. This procedure was performed separately in each ethnic group. Individuals with more than one out of 14 missing genotype were not included in the genetic risk score calculation. Although multiple SNPs included in the gene score are within the same gene, our linkage disequilibrium analyses of this cohort indicated that it is appropriate to include all 14 SNPs in the genetic risk score (adjusted for sex, age, ethnicity, glycemic status and BMI). This test was performed since a correlation between the genetic variant and environmental exposure can bias the estimates of the main genetic effect and the gene-environment interaction.

Results

Characteristics of the studied cohort. The clinical and anthropometric characteristics of the EpiDREAM study are summarized in Table 1. The mean age of participants was 52.7 years and the ethnic distribution of the cohort was 53.9% European, 18.9% Latino, 15.8% South Asian, 7.2% African, 2.9% Native American, 1.3% East Asian. The individuals in this analysis represented 17 of the original 21 countries from which recruitment took place (Supplementary Table 4). At baseline, a mean BMI of 30.2 (SD = 6.22) kg/m² and a mean BAI of 33.0 (SD = 7.49) were observed and the mean energy expenditure was 320.50 MET-minutes/week (SD = 409.20). For the present analyses, we focused on 17,423 participants at baseline and 9,228 at follow-up who had complete genotype and phenotype data. The median time between the baseline screening visit and final contact was 3.3 years. After follow-up, the mean BMI and BAI were 30.32 (SD = 5.79) and 33.78 (SD = 7.59), respectively, and the mean energy expenditure appeared to decrease slightly to 301.50 MET-minutes/week (SD = 368.04).

Effect of Physical Activity on BMI/BAI. At baseline, the quantitative MET score was significantly associated with both decreased BMI and BAI (Table 2). Congruently, the basic PA score (low – moderate – high PA) was significantly associated with lower baseline BMI and BAI. Similar associations were found between the baseline MET score and BMI and BAI at follow-up. The baseline basic PA score was also associated with decreased follow-up BMI and BAI. The baseline MET score was not associated with BMI change, or BAI change. The basic PA score was nominally associated with decreased BMI change and BAI change over follow-up.

Effect of SNPs/GRS on Physical Activity. We first investigated the association of 14 obesity predisposing SNPs and corresponding GRS on basic PA score adjusting for sex, age, ethnicity, glycemic status and BMI (Table 3). We observed a nominal association between three of these SNPs and the basic PA score: NTRK2 rs1211166, BDNF rs1401635 and NPC1 rs1805081. The association between the obesity risk GRS and the basic PA score was not significant. When adjusting for BAI rather than BMI the same three SNPs remained nominally associated with the basic PA score with a consistent direction of effect (Supplementary Table 5).
| Category                  | PA-Low   | PA-Moderate | PA-High  | All      | p-value  |
|---------------------------|----------|-------------|----------|----------|----------|
| Total at baseline N(%)    |          |             |          |          |          |
| Gender N(%)               |          |             |          |          |          |
| Male                      | 4727 (27.1%) | 10529 (60.4%) | 2151 (12.3%) | 17407 (100%) | 3.8 x 10^-13 |
| Female                    | 3229 (68.3%) | 6441 (61.2%) | 938 (45.6%) | 10608 (60.9%) |          |
| Age (years)               | 51.24 ± 12.13 | 53.17 ± 11.05 | 53.25 ± 10.97 | 52.66 ± 11.38 | 1.5 x 10^-12 |
| METS at baseline 165.15 ± 165.15 | 165.15 ± 165.15 | 165.15 ± 165.15 | 165.15 ± 165.15 | 1.8 x 10^-4 |
| BMI at baseline (kg/m²)   | 30.90 ± 7.07 (4726) | 30.90 ± 7.07 (4726) | 28.89 ± 5.17 (2153) | 30.16 ± 6.22 (17407) | 3.2 x 10^-18 |
| BMI at follow-up (kg/m²)  | 30.81 ± 6.23 (2429) | 30.35 ± 5.67 (5706) | 29.03 ± 5.11 (1070) | 30.32 ± 5.79 (9205) | 2.1 x 10^-16 |
| BMI Change (kg/m²)        | 0.08 ± 2.85 (2427) | 0.25 ± 2.41 (5704) | 0.28 ± 2.12 (1070) | 0.21 ± 2.50 (9201) | 0.22 |
| BAI at baseline           | 35.19 ± 8.08 (4716) | 32.65 ± 7.20 (10500) | 29.91 ± 6.02 (2138) | 33.00 ± 7.49 (17354) | 3.4 x 10^-13 |
| BAI at follow-up          | 35.51 ± 8.19 (2312) | 33.67 ± 7.37 (5487) | 30.48 ± 6.07 (1033) | 33.78 ± 7.59 (8832) | 2.7 x 10^-12 |
| BAI Change                | 0.09 ± 3.92 (2309) | 0.56 ± 3.65 (5482) | 0.32 ± 3.38 (1032) | 0.43 ± 3.69 (8823) | 5.0 x 10^-14 |
| Ethnic groups N(%)        |          |             |          |          |          |
| South Asian               | 1527 (32.3%) | 1122 (10.7%) | 111 (5.2%) | 2760 (15.8%) | 8.4 x 10^-13 |
| East Asian                | 34 (0.7%)  | 154 (1.5%)  | 37 (1.7%)  | 225 (1.3%)  |          |
| European                  | 1689 (35.7%) | 6091 (57.8%) | 1601 (74.4%) | 9381 (53.9%) |          |
| African                   | 362 (7.7%)  | 793 (7.5%)  | 94 (4.4%)  | 1249 (7.2%)  |          |
| Latino American           | 1002 (21.2%) | 2057 (19.5%) | 233 (10.8%) | 3292 (18.9%) |          |
| Native North American     | 113 (2.4%)  | 312 (3.0%)  | 75 (3.5%)  | 500 (2.9%)  |          |

Table 1. Baseline characteristics by physical activity status in EpiDREAM study. Notes: BMI: body mass index, BAI: body adiposity index, SD: standard deviation, N: sample size, PA: physical activity *Data are presented as mean ± SD (N)

| Outcome                        | β       | 95% CI     | P-value |
|--------------------------------|---------|------------|---------|
| Effect of MET score on BMI/BAI (adjusted for gender, age, ethnicity and glycemic status) | | | |
| Baseline BMI                   | −0.54   | −0.63 to −0.45 | 1.2 x 10^-14 |
| Baseline BAI                   | −0.59   | −0.69 to −0.49 | 8.6 x 10^-12 |
| Follow-up BMI                  | −0.45   | −0.57 to −0.34 | 1.1 x 10^-11 |
| Follow-up BAI                  | −0.59   | −0.73 to −0.45 | 3.6 x 10^-14 |
| Change BMI                     | −0.05   | −0.41 to 0.11  | 0.26 |
| Change BAI                     | 0.07    | 0.01 to 0.15   | 0.07 |

Table 2. Effect of physical activity on obesity outcomes.

None of the 14 SNPs displayed a significant association with the MET score after adjustment for sex, age, ethnicity, glycemic status and BMI (Table 3). The association between the obesity risk GRS and the MET score was also non-significant. Similar results were found when adjusting for BAI rather than BMI (Supplementary Table 5).

Of the 14 SNPs analyzed, only one (NTRK2 rs1211166) showed a nominal association with change in the basic PA score (Supplementary Table 6). None of the 14 SNPs displayed a significant association with a change in the MET score. The obesity risk GRS was not associated with change in the basic PA score or change in the MET score (Supplementary Table 6).

Effect of SNPs/GRS on BMI/BAI. At baseline, the obesity risk alleles of four SNPs were associated increased with BMI and BAI. FTO rs1421085 and CDKAL1 rs2206734 were significantly associated with greater BMI/BAI, while TNN13K rs1514176 and GIPR rs11671664 were nominally associated with increased BMI/BAI (Table 4). At baseline, the GRS was significantly associated with greater BMI and BAI.

After follow-up, three SNPs (FTO rs1421085, TNN13K, rs1514176, GIPR 11671664) and the GRS were associated with increased BMI and BAI (Table 4). CDKAL1 rs2206734 displayed a nominal association with reduced BMI and BAI change. The GRS was not associated with BMI or BAI change.
as the dependent variable. \( \text{bindicates analyses with the MET score as the dependent variable.} \)

score tertile and (2) BMI increase of 0.26 kg/m\(^2\) (\( P \text{ adjusted for sex, age, ethnicity, glycemic status and BMI.} \)

Table 3. Effect of SNPs/GRS on physical activity measures. Notes: GRS = genetic risk score, all analyses were adjusted for sex, age, ethnicity, glycemic status and BMI. \( ^a \)indicates analyses with the basic physical activity score as the dependent variable, \( ^b \)indicates analyses with the MET score as the dependent variable.

### Interaction Analyses.

Interaction tests with PA were restricted to the subset of SNPs/GRS displaying a nominal or significant association with BMI/BAI at baseline and/or at follow-up. At baseline, the MET score modified the effect of the FTO risk allele on BMI and BAI (Table 5). Each additional FTO risk allele (C) was associated with a (1) BMI increase of 0.60 kg/m\(^2\) (\( P = 1.1 \times 10^{-4} \)), BAi increase of 0.45 (\( P = 5.7 \times 10^{-4} \)) in the lowest MET score tertile and (2) BMI increase of 0.26 kg/m\(^2\) (\( P = 0.05 \)), BAi increase of 0.20 (\( P = 0.15 \)) in the highest MET score tertile. This indicates that the effect of FTO rs1421085 on BMI and BAI can be reduced by 57% and 56% respectively through PA.

The basic physical score also modified the effect of the FTO risk allele on BMI and BAI at baseline (Table 5, Fig. 1). Each additional obesity risk allele (C) was associated with a (1) BMI increase of 0.71 kg/m\(^2\) (\( P = 1.4 \times 10^{-3} \)), BAi increase of 0.62 (\( P = 2.1 \times 10^{-3} \)) in the inactive group and (2) BMI increase of 0.35 kg/m\(^2\) (\( P = 0.03 \)), BAi increase of 0.40 (\( P = 0.02 \)) in the active group. This indicates that PA is associated with a 36–51% decrease in the effect of FTO rs1421085 on obesity measures at baseline. We also conducted a sensitivity analysis to analyze the BPA x FTO rs1421085 interaction for baseline BMI in only non-diabetic patients, and found similar results (\( \beta = -0.40, 95\% \text{ CI} = -0.62 \text{ to} -0.19, P = 2.4 \times 10^{-4} \)).

The basic PA score interacted with FTO rs1421085 in modulating BMI and BAI at follow-up. Each additional obesity risk allele (C) was associated with a (1) BMI increase of 0.72 kg/m\(^2\) (\( P = 1.1 \times 10^{-3} \)), BAi increase of 0.75 (\( P = 3.4 \times 10^{-4} \)) in the inactive group and (2) BMI increase of 0.19 kg/m\(^2\) (\( P = 0.39 \)), BAi increase of 0.19 (\( P = 0.44 \)) in the active group. This corresponds to a 74–75% decrease in the effect of FTO rs1421085 on BMI/BAI at follow-up.

No significant interactions were observed between TNNI3K rs1514176, CDKAL1 rs2206734, GIPR rs11671664 or the obesity risk GRS and the basic PA/MET score on BMI/BAI at baseline, follow-up or change.

Given that Ahmad et al.\(^{13} \) reported the FTO x PA interaction to be 10-fold larger in North American compared to European cohorts, we also analyzed a 3-way interaction (FTO x PA x North American residence) among the subgroup of European participants to follow-up this finding. Although the 3-way interaction was not significant (\( \beta = -0.01, 95\% \text{ CI} = -0.03 \text{ to} 2.8 \times 10^{-3}, P = 0.10 \)), we acknowledge that our statistical power to detect 3-way interactions in this subgroup was limited and this association warrants further investigation.

### Discussion

We observed significant GEI between FTO rs1421085 and PA at baseline and at follow-up in an international multiethnic population using both basic and quantitative assessments of PA. Although the interactions were nominally significant, the Bonferroni correction we applied is overly conservative, particularly when testing highly correlated outcomes\(^3\) such as BMI/BAI, basic PA/MET score, and we have confidence in our results for several reasons. First, the power to detect GEI at the nominal level of association (\( P < 0.05 \)) is adequate at baseline, although our power estimations for the Bonferroni corrected threshold are moderate (see Supplementary Figures 3–6). Second, twin studies have shown that PA can substantially reduce the influence of genetic factors on BMI in adults\(^2\). Third, the interaction between FTO and PA has been demonstrated in several cross-sectional studies, and currently represents the most robust example of GEI in the field of genetic epidemiology\(^8–11,18\). Fourth, there is a plausible underlying biological process to substantiate this association. FTO is a nucleic acid demethylase and FTO intron 1 variation is associated with different methylation profiles and BMI variance\(^3\). Since methylation of DNA is sensitive to environmental changes (e.g. PA and diet) there is a strong biological rationale to identify GEI with FTO as previously reported\(^6,35\). Two studies have shown that PA can change the methylation and mRNA expression pattern of genes, including FTO, in both muscle and adipose tissue\(^24,26\). A more recent analysis demonstrated that variation at the FTO locus represses mitochondrial thermogenesis in adipocyte precursor cells and causes a shift from energy dissipating beige (brite) to energy-storing white adipocytes, which is accompanied by increased lipid storage and weight gain\(^17\). Exercise studies in humans and mouse models indicate that exercise...
training increases the expression of the brown adipocyte marker uncoupling protein (UCP1) in both visceral and subcutaneous white adipose tissue (WAT)\textsuperscript{38}. These changes are associated with increases in brown-like adipocytes (browning or beiging), particularly in subcutaneous WAT\textsuperscript{38}. The combined role of FTO and physical activity in obesity and adipocyte browning, in conjunction with epigenetic mechanisms, strengthen the biological rationale and confidence in the statistical interaction\textsuperscript{39}.

Given the growing consensus that food intake may be the main driver of the obesity epidemic\textsuperscript{40}, it is important to note that both PA measures displayed significant associations with both adiposity measures at baseline and at follow-up. This indicates that PA can influence obesity, despite the broad range of lifestyles among the participants. The value of PA for managing obesity has been recognized in a recent analysis of the National Health and Nutrition Examination Survey (NHANES) cohort from 1988–2010, which found that PA had a larger impact on BMI and waist circumference trends than calorie intake\textsuperscript{41}. Our cross sectional analyses indicate that one hour of jogging or swimming (8.0 MET activities) per week was associated with approximately a 0.5 kg/m\textsuperscript{2} decrease in BMI. Together, these data challenge the idea of attributing the obesity epidemic mainly to excessive caloric intake\textsuperscript{42} and support the universal value of PA to maintain a healthy body weight\textsuperscript{43}.

Although the MET score provides a more comprehensive assessment of PA participation, only 11 015 (63\%) participants completed the assessment of the MET score, compared to the 17 407 (99\%) participants who completed the basic PA score. The loss of power induced by the smaller sample size may have been compensated for by the added precision of the MET score. Simulations of GEI have shown that a sample of about 2 000 participants with precisely measured environmental exposure and outcome data are needed to detect a GEI of large magnitude (a doubling of the genetic risk estimate in the exposed group compared to the unexposed group) with reasonable power (95\% power, \( P = 1 \times 10^{-4} \))\textsuperscript{43}. With less precise measurement of environmental exposure, the sample size requirement can increase to 100 000 participants to detect the same interaction with comparable power\textsuperscript{44,45}. Using a brief PA assessment may be the best compromise to balance the sample size requirements with the need for sufficient precision, as recently suggested by Peters et al\textsuperscript{44}.

Measuring body fat content data is less feasible in large sample sizes for GEI studies. Only two small GEI studies (\( N < 800 \)) have used direct body fat content measures (DEXA and underwater weighing)\textsuperscript{46,47} and the meta-analysis
| Interaction terms                  | SNP Main Effect | SNP Interaction |
|-----------------------------------|-----------------|-----------------|
|                                   | β               | 95% CI          | P               | β               | 95% CI          | P               |
| **Baseline Interaction Tests**    |                 |                 |                 |                 |                 |                 |
| **Outcome: BMI baseline**         |                 |                 |                 |                 |                 |                 |
| **FTO rs1421085 x MET score**     | 1.26            | 0.47–2.05       | $1.9 \times 10^{-5}$ | $-0.14$         | $-0.26$–$-0.01$ | 0.03            |
| Subgroup analysis                 |                 |                 |                 |                 |                 |                 |
| Low MET tertile                   |                 |                 |                 |                 |                 |                 |
| Middle MET tertile                |                 |                 |                 |                 |                 |                 |
| High MET tertile                  |                 |                 |                 |                 |                 |                 |
| **TNNI3K rs1514176 x MET score**  | 0.38            | $-0.38$–$-1.15$ | 0.33            | $-0.02$         | $-0.14$–$-0.10$ | 0.72            |
| **CDKAL1 rs2206734 x MET score**  | 0.27            | $-0.69$–$-1.23$ | 0.58            | $-3.4 \times 10^{-3}$ | $-0.16$–$-0.15$ | 0.97            |
| **GIPR rs11671664 x MET score**   | 1.48            | 0.27–2.68       | 0.02            | $-0.18$         | $-0.37$–$-0.01$ | 0.07            |
| **GRS x MET score**               | 0.60            | 0.37–0.84       | $4.3 \times 10^{-3}$ | $-0.07$         | $-0.11$–$-0.04$ | 0.11            |
| **FTO rs1421085 x BPA score**     | 1.08            | 0.69–1.47       | $5.4 \times 10^{-4}$ | $-0.34$         | $-0.54$–$-0.14$ | $8.5 \times 10^{-4}$ |
| Subgroup analysis                 |                 |                 |                 |                 |                 |                 |
| Low PA                            |                 |                 |                 |                 |                 |                 |
| Moderate PA                       |                 |                 |                 |                 |                 |                 |
| High PA                           |                 |                 |                 |                 |                 |                 |
| **TNNI3K rs1514176 x BPA score**  | $-0.17$         | $-0.55$–$-0.21$ | 0.38            | 0.20            | 0.01–0.40       | 0.04            |
| **CDKAL1 rs2206734 x BPA score**  | 0.65            | 0.18–1.13       | $6.8 \times 10^{-3}$ | $-0.21$         | $-0.45$–$-0.03$ | 0.09            |
| **GIPR rs11671664 x BPA score**   | 0.48            | $-0.13$–$-1.09$ | 0.12            | $-0.10$         | $-0.41$–$-0.21$ | 0.52            |
| **GRS x BPA score**               | 0.22            | 0.11–0.34       | $1.2 \times 10^{-4}$ | $-0.05$         | $-0.11$–$-0.01$ | 0.09            |
| **Outcome: BAI baseline**         |                 |                 |                 |                 |                 |                 |
| **FTO rs1421085 x MET score**     | 1.21            | 0.36–2.07       | $5.4 \times 10^{-5}$ | $-0.14$         | $-0.28$–$-0.01$ | 0.04            |
| Subgroup analysis                 |                 |                 |                 |                 |                 |                 |
| Low MET tertile                   |                 |                 |                 |                 |                 |                 |
| Middle MET tertile                |                 |                 |                 |                 |                 |                 |
| High MET tertile                  |                 |                 |                 |                 |                 |                 |
| **TNNI3K rs1514176 x MET score**  | 0.34            | $-0.48$–$-1.17$ | 0.42            | $-0.02$         | $-0.15$–$-0.11$ | 0.73            |
| **CDKAL1 rs2206734 x MET score**  | $-0.10$         | $-1.13$–$-0.94$ | 0.85            | 0.06            | $-0.11$–$-0.23$ | 0.45            |
| **GIPR rs11671664 x MET score**   | 1.06            | $-0.24$–$-2.35$ | 0.11            | $-0.12$         | $-0.33$–$-0.09$ | 0.25            |
| **GRS x MET score**               | 0.47            | 0.21–0.72       | $2.9 \times 10^{-4}$ | $-0.06$         | $-0.10$–$-0.02$ | 0.07            |
| **FTO rs1421085 x BPA score**     | 0.95            | 0.53–1.37       | $1.0 \times 10^{-4}$ | $-0.31$         | $-0.53$–$-0.10$ | $4.4 \times 10^{-3}$ |
| Subgroup analysis                 |                 |                 |                 |                 |                 |                 |
| Low PA                            |                 |                 |                 |                 |                 |                 |
| Moderate PA                       |                 |                 |                 |                 |                 |                 |
| High PA                           |                 |                 |                 |                 |                 |                 |
| **TNNI3K rs1514176 x BPA score**  | $-0.21$         | $-0.62$–$-0.20$ | 0.32            | 0.19            | $-0.02$–$-0.40$ | 0.08            |
| **CDKAL1 rs2206734 x BPA score**  | 0.57            | 0.06–1.08       | 0.03            | $-0.14$         | $-0.40$–$-0.13$ | 0.31            |
| **GIPR rs11671664 x BPA score**   | 0.34            | $-0.32$–$-1.00$ | 0.31            | $-0.06$         | $-0.40$–$-0.28$ | 0.72            |
| **GRS x BPA score**               | 0.19            | 0.07–0.31       | $2.7 \times 10^{-5}$ | $-0.04$         | $-0.10$–$-0.03$ | 0.27            |
| **Follow-up Interaction Tests**   |                 |                 |                 |                 |                 |                 |
| **Outcome: BMI follow-up**        |                 |                 |                 |                 |                 |                 |
| **FTO rs1421085 x MET score**     | 1.33            | 0.29–2.37       | 0.01            | $-0.16$         | $-0.33$–$-0.01$ | 0.06            |
| **TNNI3K rs1514176 x MET score**  | 0.63            | $-0.38$–$-1.64$ | 0.22            | $-0.06$         | $-0.22$–$-0.10$ | 0.48            |
| **GIPR rs11671664 x MET score**   | 0.25            | $-1.41$–$-1.91$ | 0.77            | 0.01            | $-0.26$–$-0.27$ | 0.96            |
| **GRS x MET score**               | 0.47            | 0.16–0.79       | $3.2 \times 10^{-3}$ | $-0.06$         | $-0.11$–$-0.01$ | 0.18            |
| **FTO rs1421085 x BPA score**     | 1.10            | 0.59–1.61       | $2.3 \times 10^{-5}$ | $-0.36$         | $-0.62$–$-0.10$ | $6.8 \times 10^{-3}$ |
| Subgroup analysis                 |                 |                 |                 |                 |                 |                 |
| Low PA                            |                 |                 |                 |                 |                 |                 |
| Moderate PA                       |                 |                 |                 |                 |                 |                 |
| **Continued**                     |                 |                 |                 |                 |                 |                 |
Interaction terms | SNP Main Effect | SNP Interaction
---|---|---
| β | 95% CI | P | β | 95% CI | P |
TNNI3K rs1514176 x BPA score | High PA | 0.19 kg/m² | −0.25–0.64 | 0.39 |
GIPR rs11671664 x BPA score | 0.56 | −0.24–1.36 | 0.17 | −0.16 | −0.57–0.24 | 0.43 |
GRS x BPA score | 0.15 | −7.3 × 10⁻⁴–0.30 | 0.05 | −0.29 | −0.11–0.05 | 0.47 |

Outcome: BAI follow-up

| Interaction terms | SNP Main Effect | SNP Interaction |
---|---|---|
FTO rs1421085 x MET score | 1.91 | 0.69–3.14 | 2.3 × 10⁻³ | −0.24 | −0.43–0.04 | 0.17 |
TNNI3K rs1514176 x MET score | 0.57 | −0.61–1.75 | 0.34 | −0.05 | −0.23–0.14 | 0.64 |
GIPR rs11671664 x MET score | −0.09 | −2.03–1.85 | 0.93 | 0.08 | −0.24–0.39 | 0.63 |
GRS x MET score | 0.57 | 0.20–0.93 | 2.7 × 10⁻³ | −0.08 | −0.14–0.02 | 0.07 |
FTO rs1421085 x BPA score | 1.22 | 0.60–1.84 | 1.2 × 10⁻³ | −0.40 | −0.71–0.08 | 0.01 |
Subgroup analysis | Low PA | 0.75 | −0.34–1.17 | 3.4 × 10⁻⁴ |
Moderate PA | 0.43 | 0.18–0.68 | 6.9 × 10⁻⁴ |
High PA | 0.19 | −0.29–0.67 | 0.44 |
TNNI3K rs1514176 x BPA score | 0.18 | −0.42–0.78 | 0.56 | 0.01 | −0.30–0.32 | 0.96 |
GIPR rs11671664 x BPA score | 0.75 | −0.22–1.72 | 0.13 | −0.21 | −0.70–0.29 | 0.41 |
GRS x BPA score | 0.21 | 0.03–0.39 | 0.02 | −0.06 | −0.16–0.03 | 0.18 |

Outcome: BMI Change

| Interaction terms | SNP Main Effect | SNP Interaction |
---|---|---|
CDKAL1 rs2206734 x BPA score | −0.18 | −0.45–0.11 | 0.22 | 0.02 | −0.13–0.17 | 0.81 |

Outcome: BAI Change

| Interaction terms | SNP Main Effect | SNP Interaction |
---|---|---|
CDKAL1 rs2206734 x BPA score | −0.20 | −0.64–0.23 | 0.36 | 0.01 | −0.21–0.24 | 0.92 |

Table 5. Interaction analyses between physical activity measures and obesity predisposing SNPs/GRS. Notes: BPA score = basic physical activity score. GRS = Genetic risk Score.

Figure 1. Mean baseline BMI values stratified by physical activity level (PA) and FTO rs1421085 genotype.

of the PA x FTO interaction which analyzed bioelectrical impedance was over 95% Europeans16. To our knowledge, this is the first large-scale study to report an interaction between PA and FTO rs1421085 using the BAI. Since assessing body fat content is financially prohibitive in large-scale analyses, the BAI may be an acceptable method to complement the widely used BMI measure to assess adiposity in GEI studies.

Assessing the impact of the 14 SNPs on PA identified some interesting nominal associations. Despite being an obesity predisposing gene, BDNF may be related to increased PA levels in our cohort. While this appears contradictory, the loss of one functional copy of BDNF gene has been associated with Mendelian obesity, cognitive impairment and hyperactivity in both humans47 and rodents48. It has been proposed that increased PA may be a behavioral response to compensate for their proclivity to gain weight49. Alternatively, these SNPs have been hypothesized to be a remnant of our hunter-gatherer past, and may have been positively selected since they promote an “active-foraging” phenotype that induces a preference for energy-dense foods and the physical disposition to attain those foods49. Together, this food seeking behaviour and overactive disposition represent a “Peppy-Thrifty
Genotype.” In contrast, NPC1 variation appears to contribute to sedentary behavior and may complement the thrifty genotype hypothesis. Genetic variants increasing food intake and decreasing PA may have been positively selected based on their parallel effects on energy balance. The contrasting effects of these obesity predisposing SNPs on PA may account for the challenges in substantiating the lazy thrifty genotype hypothesis. Although there appears to be some shared genetic correlation between obesity and PA, additional studies are necessary to confirm our nominally significant associations and to clarify the impact of additional obesity predisposing genes on PA.

Limitations of this study include the multi-ethnic composition of the EpiDREAM cohort may have added significant heterogeneity in the analyses, especially as important PA differences are observed in different ethnic backgrounds. Although, the MET score was calculated with more objective criteria of PA participation, recalling participation in 41 different activities introduces a source of error and or recall bias. Since most of the obesity predisposing SNPs selected in the study were originally identified in European populations, they may not be ideal proxies for the causal SNP in other ethnic groups. We are aware that the 14 SNPs analyzed only represent a subset of the current list of validated obesity SNPs. Lastly, the EpiDREAM population (participants identified for hyper-glycemia risk) is not representative of the general population, and the participants missing from the follow-up analysis may have created a systematic bias in our sample. However, this bias may not have influenced our results since no significant differences in BMI (P = 0.25) or BAI (P = 0.21) were observed among those who completed follow-up and those that did not.

Strengths of this analysis include the precise MET score, complementing the BMI with the recently developed BAI, the prospective cohort design and the multi-ethnic sample.

In summary, we identified an interaction between the FTO SNP rs1421085 and PA in a prospective cohort of six ethnic groups from 17 countries. While this has been demonstrated previously, this is the first study to analyze this interaction prospectively using a quantitative measure of PA while comparing the recently developed BAI and BMI. Analyzing the impact of obesity predisposing SNPs on PA revealed novel associations, although further study is needed to confirm these effects. These findings suggest that obesity prevention programs emphasizing vigorous PA for genetically at risk subgroups may be a valuable contribution to the global fight against obesity.

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