The Association Between Long-Term Exposure to Particulate Matter and Incidence of Hypertension Among Chinese Elderly: A Retrospective Cohort Study

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Background and Objectives: Studies that investigate the links between particulate matter ≤2.5 µm (PM$_{2.5}$) and hypertension among the elderly population, especially those including aged over 80 years, are limited. Therefore, we aimed to examine the association between PM$_{2.5}$ exposure and the risk of hypertension incidence among Chinese elderly.

Methods: This prospective cohort study used 2008, 2011, 2014, and 2018 wave data from a public database, the Chinese Longitudinal Healthy Longevity Survey, a national survey investigating the health of those aged over 65 years in China. We enrolled cohort participants who were free of hypertension at baseline (2008) from 706 counties (districts) and followed up in the 2011, 2014, and 2018 survey waves. The annual PM$_{2.5}$ concentration of 706 counties (districts) units was derived from the Atmospheric Composition Analysis Group database as the exposure variable, and exposure to PM$_{2.5}$ was defined as 1-year average of PM$_{2.5}$ concentration before hypertension event occurrence or last interview (only for censoring). A Cox proportional hazards model with penalized spline was used to examine the non-linear association between PM$_{2.5}$ concentration and hypertension risk. A random-effects Cox proportional hazards model was built to explore the relationship between each 1 µg/m$^3$, 10 µg/m$^3$ and quartile increment in PM$_{2.5}$ concentration and hypertension incidence after adjusting for confounding variables. The modification effects of the different characteristics of the respondents were also explored.

Results: A total of 7,432 participants aged 65–116 years were enrolled at baseline. The median of PM$_{2.5}$ exposure concentration of all the participants was 52.7 (inter-quartile range, IQR = 29.1) µg/m$^3$. Overall, the non-linear association between PM$_{2.5}$ and hypertension incidence risk indicated that there was no safe threshold for PM$_{2.5}$ exposure. The higher PM$_{2.5}$ exposure, the greater risk for hypertension incidence. Each 1 µg/m$^3$ (adjusted hazard ratio (AHR): 1.01; 95% CI: 1.01–1.02) and 10 µg/m$^3$ (AHR: 1.12; 95% CI: 1.09–1.16) increments in PM$_{2.5}$, were associated with the incidence of hypertension after adjusting for potential confounding variables. Compared to first quartile (Q1) exposure, the adjusted HRs of hypertension incidence for the Q2, Q3 and
Q4 exposure of PM$_{2.5}$ were 1.31 (95% CI: 1.13–1.51), 1.35 (95% CI: 1.15–1.60), and 1.83 (95% CI: 1.53–2.17), respectively. The effects appear to be stronger among those without a pension, living in a rural setting, and located in central/western regions.

**Conclusion:** We found no safe threshold for PM$_{2.5}$ exposure related to hypertension risk, and more rigorous approaches for PM$_{2.5}$ control were needed. The elderly without a pension, living in rural and setting in the central/western regions may be more vulnerable to the effects of PM$_{2.5}$ exposure.

**Keywords:** China, hypertension, elderly, particulate matter, cohort study

**INTRODUCTION**

Hypertension is the most prevalent chronic disease among the elderly (1) and has led to adverse cardiovascular diseases (CVDs), such as hypertensive heart disease and stroke, and even death (2). Worldwide, as the elderly population continues to grow, the disease burden attributed to hypertension and its complications are increasing (3, 4) so it is imperative to identify risk factors for hypertension and promote prevention among the elderly.

In addition to well-established associations with lifestyle and heredity, it is thought that environmental pollutants also contribute to the occurrence of hypertension (5, 6). In particular, the inhalation of air containing fine particulate matter (PM), especially sizes $\leq$2.5 um (PM$_{2.5}$) has been reported as a probable antecedent driver of the increase in blood pressure and incidence of hypertension (7–12). Overall, although the association between PM$_{2.5}$ exposure and hypertension has been reported in several studies, gaps in knowledge remain. For example, elderly people may be more susceptible to PM$_{2.5}$, due to higher rates of CVDs and given the decline in organ function associated with age (13, 14). However, relevant studies rarely focus on the elderly, especially those aged over 80 years (15–19).

Furthermore, inconsistent results have been reported in previous studies. For instance, based on a cross-sectional study involving 27,752 Taipei City residents (mean age 74.8 years), Chen et al. found that PM$_{2.5}$ exposure was not significant associated with diastolic blood pressure (DBP) and none of the air pollutants were associated with changes in systolic blood pressure (SBP) (16). In addition, cross-sectional study designs and small-scale sample sizes limit the generalizability of the results to the elderly population (17, 19–21).

Lastly, many studies have been conducted in regions and countries, such as Taiwan and the USA (16, 21), with good air quality defined by PM$_{2.5}$ concentrations lower than the World Health Organization recommendations (10 $\mu$g/m$^3$). It has been suggested that the conclusions drawn from studies conducted in other regions might not be suitable for generalization to China (11, 22). China has some of the heaviest air pollution in the world with an average annual population-weighted PM$_{2.5}$ exposure of 52.7 $\mu$g/m$^3$ in 2017 (23). More importantly, PM$_{2.5}$ has been proven to be a modifiable factor that contributes to cardiovascular morbidity and mortality (24). Thus, considering the research gaps and public health significance, it is necessary to study and verify the effects of PM$_{2.5}$ on hypertension among the elderly to provide more effective interventions for the elderly and medical resource allocation.

In the present prospective cohort study based on the national representative Chinese Longitudinal Healthy Longevity Survey (CLHLS), we aimed to examine the association between PM$_{2.5}$ exposure and the risk of hypertension incidence among Chinese elderly. Secondarily, we aimed to further explore the modifying effects of PM$_{2.5}$ exposure on hypertension incidence and identify vulnerable sub-populations.

**MATERIALS AND METHODS**

**Study Population**

This was a prospective cohort study that selected data from the 2008, 2011, 2014, and 2018 waves of the CLHLS. The CLHLS is a nationwide survey that covers 23 of 31 provinces, municipalities, and autonomous regions in mainland China (the remaining provinces are not surveyed because of their low level of population density), which includes 85% of the total population of the country (25). The CLHLS was established in 1998 with enrollment of the elderly population (aged 65 years and older) and traced them in 1998, 2000, 2002, 2005, 2008, 2011, and 2018. In the 2008 wave, the CLHLS added data collection on individuals’ community-level information, including economic development, the natural environment, and environmental pollution. The CLHLS study was approved by the Institutional Review Board of the Duke University Health System and all participants sign an informed consent form. A detailed introduction to the CLHLS has been published previously (26).

The cohort for the present study was derived from 10 years of follow-up (2008–2018) from the 2008 baseline when the availability of county (district) unit level address information became available. The cohort comprised 7,432 participants from 706 county (district) units aged 65–116 years. The participants were all free of hypertension (normal blood pressure at the time of the survey and self-reported as not being diagnosed with hypertension by a physician) and had complete demographic characteristic information at baseline in 2008. They were interviewed in 2011, 2014, and 2018. A flow chart of the study population selection is shown in Figure 1. The follow-up duration was reported as person-years calculated by using days divided by 365 from...
the date of study enrollment to the date of last interview, death, or hypertension incidence (whichever came first). The outcome of the event was considered as the occurrence of hypertension, while other conditions such as death, lost follow-up, and not been identified as hypertension were censored data.

Assessment of Hypertension
As part of the CLHLS study, blood pressure values for each participant were collected by trained investigators using a mercurial sphygmomanometer (upper arm type; Yuyue, Jiangsu, China) at baseline and at every follow-up interview. Their blood pressure was measured two times after a five or more min rest period. Specifically, for bedbound participants, blood pressure measurements were obtained in a recumbent position. SBP and DBP values were calculated as the average of two repeated measurements. According to the Chinese Guidelines for Prevention and Treatment of Hypertension (27), we defined hypertension by an SBP ≥ 140 mmHg, DBP ≥ 90 mmHg, or normal blood pressure with a self-reported hypertension diagnoses previously made by a grade II or III hospital. The definition of hypertension was supported by prior study using same database (28).

Assessment of PM$_{2.5}$ Exposure
Due to the privacy protection, the information of participant’s home address was deleted in the CLHLS. Thus, we identified the 706 residence name [i.e., county (district) from community environment questionnaire in this study], detailed strategy was published in a prior study (29). The annual average residential PM$_{2.5}$ concentration from 706 residential county (district) units of the 7,432 participants was collected through an open database built by the Atmospheric Composition Analysis Group from the University of Washington (https://sites.wustl.edu/acag/) (30). The database collects ground-level PM2.5 measurements were obtained from http://beijingair.sinaapp.com/ over mainland China. These data are captured by individuals from instantaneous data records on the website of the Chinese EPA. The PM$_{2.5}$ was evaluated from satellite observations with 0.01 spatial resolution (1.1 × 1.1 km). The resultant PM$_{2.5}$ estimates were highly convergent ($R^2 = 0.81$) with out-of-sample cross-validated PM$_{2.5}$ concentrations from monitors (31). This database has been widely used in previous studies (22, 32).

Referring to prior studies (33, 34), the PM$_{2.5}$ exposure was defined as 1-year average before the hypertension event occurrence or the last interview (only for censored data). This exposure time was selected because it has the strongest
hazard risk for hypertension (Supplementary Table 1). In our study, only 116 (1.5%) of participants moved to another address (county/district level), which was similar to prior study (22). Thus, we hypothesized that the PM$_{2.5}$ exposure of the population was stable, we also performed sensitivity analysis through excluding participants who changed address (Supplementary Table 3).

**Potential Confounding Variables**

According to previous studies, several demographic, lifestyle, and health status variables at baseline were considered potential confounding factors and adjusted for, including sex (male or female), residence (rural or urban areas), age (65–80 or >80 years), regions (eastern China or central/western China), living arrangement, pension (yes or no), educational attainment (0 or >0 years), marital status (separated/divorced/never married, widowed, or married), body mass index (BMI) (<18.5, 18.5–23.9, or >23.9), smoking at the present (yes or no), drinking at the present (yes or no), exercising at the present (yes or no), history of diabetes (yes or no), heart disease (yes or no) and function disability. Among them, function disability was measured by Activities of daily living (ADL) scale (35). The scale has six items including the fundamental skills: bathing, dressing, eating, toileting, continence, and transferring. If a respondent is able to perform an activity, he/she gets score 1, and if he/she is limited to do and unable to do so, will get score 2 and 3, respectively. The more scores of the individual, indicates poor ADL ability. According to the cutoff value (six points), the function disability was set as binary variable in the analysis (yes or no).

The per capita gross domestic product (GDP) (≥median or <median) and proportion of secondary industry (≥median or <median) in the city level at baseline were also considered as confounding variables in the analysis.

**Statistics Analysis**

A descriptive analysis was conducted for all the variables. Continuous variables are expressed as mean (±SD) or median (interquartile range). Categorical variables are reported as numbers and percentages. We conducted a statistical map to describe the spatial distribution of PM$_{2.5}$ concentration in 706 (counties or districts) unit-level residences at baseline. To identify the potential linear or non-linear relationship between PM$_{2.5}$ exposure and hazard ratios (HRs) of hypertension incidence, we built a Cox proportional hazards model with penalized splines and different degrees of freedom based on the minimum value of Akaike information criteria (AIC) to visualize the exposure-response relationship between PM$_{2.5}$ and hypertension. Because the multiple-level structure of CLHLS showed the clustering of participants at the study level (counties or districts) units, we used a random-effects Cox proportional hazards model for the clustering level to explore the relationship between PM$_{2.5}$ exposure and hypertension incidence among the elderly. We classified PM$_{2.5}$ concentration in increments of 1 and 10 µg/m$^3$. The estimation of the risk of hypertension incidence based on 1 and 10 µg/m$^3$ increments in PM$_{2.5}$ concentration were conducted using the random-effects Cox proportional hazards model with PM$_{2.5}$ exposure as a continuous variable. According to quartiles, PM$_{2.5}$ was categorized into four groups (Q1: ≤P$_{25}$, Q2: (P$_{25}$–P$_{75}$), Q3: (P$_{50}$–P$_{75}$), and Q4: >P$_{75}$), and the first group (Q1) was coded as a reference to examine the association between exposure and hypertension incidence. The crude model only included the PM$_{2.5}$ variable. Model 1 further added age, sex, educational attainment, pension, living arrangement, marital status, regions, residence, smoking at the present, BMI, drinking at present, exercising at the present, self-reported diabetes, function disability, GDP per capita, the proportion of secondary industry and heart disease. Analysis was also performed to examine the linear trend between the PM$_{2.5}$ quartile and hypertension incidence. The HRs and 95% confidence intervals (CIs) were calculated to evaluate the effects of PM$_{2.5}$ exposure on hypertension after adjusting for potential variables.

We also conducted subgroup analysis to evaluate whether the effect of PM$_{2.5}$ exposure on hypertension incidence differed by sex, age, educational attainment, living arrangement, residence, pension, marital status, regions, smoking at the present, drinking alcohol at the present, exercising at the present, BMI, self-reported diabetes, function disability, GDP per capita, the proportion of secondary industry and heart disease after adjusting for related covariates. Referring to a prior study (36), a 2-sample test for assessing statistically significant differences in the estimated HR within each subgroup was performed using the point estimate and standard error (SE) in this study (see Formula (1)).

\[
Z = \frac{Q_1 - Q_2}{\sqrt{(SE_1)^2 + (SE_2)^2}}
\]

In the equation, $Q_1$ and $Q_2$ are the estimated hazard ratios for each stratum, respectively. SE$_1$ and SE$_2$ are the standard errors for each stratum, respectively.

The map of China was derived from National Geomatics Center of China (http://www.ngcc.cn/ngcc/), the statistics map was generated by ArcGIS Geospatial Analyst module v10.6 (ESRI, Redlands, CA, USA). All analyses were performed using R statistical software (R 4.0.5, R Foundation for Statistical Computing, Vienna, Austria). A two-sided P-value < 0.05 was used to assess statistical significance.

**RESULTS**

At baseline, 7,432 participants were enrolled in the study. During the 10-year follow-up, the total number of person-years was 24,222 and the incidence of hypertension was 8.5 per 100 person-years. The demographic characteristics, lifestyles, and health status of the participants from the CLHLS at baseline are presented in Table 1. The mean age of the participants was 87.7 (±11.5) years. Of the elderly, 56.4% were female and 62.6% were uneducated. Over half of them had no pension. Of the participants, 80.8% lived in rural areas and 83.7% lived with family members. More participants came from western/central China than from eastern China. A total of 17.6 and 18.4% of
TABLE 1 | The demographic characteristics of the participants from the CLHLS at baseline \( n = 7,432 \). Data are number (%) of participants except PM\(_{2.5}\) concentration [median (IQR)].

| Characteristic of participants in the study | Entire cohort | \( \text{PM}_{2.5} \) |
|--------------------------------------------|---------------|-----------------|
|                                            | Q1            | Q2              | Q3              | Q4              |
| Gender                                     |               |                 |                 |                 |
| Female                                     | 4,193 (56.4)  | 992 (63.4)      | 1,047 (56.2)    | 1,079 (58.2)    | 1,075 (57.9)    |
| Male                                       | 3,239 (43.6)  | 866 (46.6)      | 816 (43.8)      | 774 (41.8)      | 783 (42.1)      |
| Age (year)                                 |               |                 |                 |                 |                 |
| 65–80                                      | 2,141 (28.8)  | 644 (34.7)      | 548 (29.4)      | 500 (27.0)      | 449 (24.2)      |
| >80                                        | 5,291 (71.2)  | 1,214 (66.3)    | 1,315 (70.6)    | 1,353 (73.0)    | 1,409 (75.8)    |
| Education level (year)                     |               |                 |                 |                 |                 |
| 0                                          | 4,654 (62.6)  | 1,120 (60.3)    | 1,096 (58.8)    | 1,157 (62.4)    | 1,281 (68.9)    |
| >0                                         | 2,778 (37.4)  | 738 (39.7)      | 767 (41.2)      | 696 (37.6)      | 577 (31.1)      |
| Pension                                    |               |                 |                 |                 |                 |
| No                                         | 6,162 (82.9)  | 1,573 (84.7)    | 1,536 (82.4)    | 1,486 (80.2)    | 1,567 (84.3)    |
| Yes                                        | 1,270 (17.1)  | 285 (15.3)      | 327 (17.6)      | 367 (19.8)      | 291 (15.7)      |
| Residence                                  |               |                 |                 |                 |                 |
| Rural                                      | 6,004 (80.8)  | 1,537 (82.7)    | 1,589 (85.3)    | 1,436 (77.5)    | 1,442 (77.6)    |
| Urban                                      | 1,428 (19.2)  | 321 (17.3)      | 274 (14.7)      | 417 (22.5)      | 416 (22.4)      |
| Living arrangement                         |               |                 |                 |                 |                 |
| Nursing institution/alone                  | 1,210 (16.3)  | 322 (17.3)      | 354 (19.0)      | 260 (14.0)      | 274 (14.7)      |
| Living in home with family member(s)       | 6,222 (83.7)  | 1,536 (82.7)    | 1,509 (81.0)    | 1,593 (86.0)    | 1,584 (85.3)    |
| Regions                                    |               |                 |                 |                 |                 |
| Western/central China                      | 4,477 (60.2)  | 1,064 (57.3)    | 1,188 (63.8)    | 856 (46.2)      | 1,369 (73.7)    |
| China                                      | 2,955 (39.8)  | 794 (42.7)      | 675 (36.2)      | 997 (53.8)      | 489 (26.3)      |
| Current marital status                     |               |                 |                 |                 |                 |
| Widowed/separated/divorced/never married   | 5,112 (66.8)  | 1,179 (63.5)    | 1,254 (67.3)    | 1,327 (71.6)    | 1,352 (72.8)    |
| Married                                    | 2,320 (31.2)  | 679 (36.5)      | 609 (32.7)      | 526 (28.4)      | 506 (27.2)      |
| Smoking at the present                     |               |                 |                 |                 |                 |
| No                                         | 6,124 (82.4)  | 1,557 (83.8)    | 1,543 (82.8)    | 1,497 (80.8)    | 1,527 (82.2)    |
| Yes                                        | 1,308 (17.6)  | 301 (16.2)      | 320 (17.2)      | 356 (19.2)      | 331 (17.8)      |
| Drink alcohol at the present               |               |                 |                 |                 |                 |
| No                                         | 6,063 (81.6)  | 1,548 (83.3)    | 1,504 (80.7)    | 1,508 (81.3)    | 1,505 (81.0)    |
| Yes                                        | 1,369 (18.4)  | 310 (16.7)      | 359 (19.3)      | 347 (18.7)      | 353 (19.0)      |
| Exercising at the present                  |               |                 |                 |                 |                 |
| No                                         | 5,397 (72.6)  | 1,330 (71.6)    | 1,328 (71.3)    | 1,380 (74.5)    | 1,359 (73.1)    |
| Yes                                        | 2,035 (27.4)  | 528 (28.4)      | 535 (28.7)      | 473 (25.5)      | 499 (26.9)      |
| Diabetes                                   |               |                 |                 |                 |                 |
| No                                         | 7,304 (98.3)  | 1,838 (98.9)    | 1,829 (98.2)    | 1,811 (97.7)    | 1,826 (98.3)    |
| Yes                                        | 128 (1.7)     | 20.0 (1.1)      | 34.0 (1.8)      | 42.0 (2.3)      | 32.0 (1.7)      |
| Heart disease                              |               |                 |                 |                 |                 |
| No                                         | 7,002 (94.2)  | 1,746 (94.0)    | 1,793 (96.2)    | 1,737 (93.7)    | 1,726 (92.9)    |
| Yes                                        | 430 (5.8)     | 112 (6.0)       | 70.0 (3.8)      | 116 (6.3)       | 132 (7.1)       |
them had smoking and drinking habits, respectively. A total of 27.4% of the participants reported that they had a habit of exercising at the present. Only 1.7% and 5.8% of the patients reported having diabetes and heart disease, respectively. A total of 52.7% of the participants had a BMI between 18.5 and 23.9. Most of them without function disability. In our study, the range of 52.7% of the participants had a BMI between 18.5 and 23.9.

The incidence of hypertension in the first quartile, Q2, Q3, and Q4 were 9.3, 8.8, 7.3, and 8.5 per 100 person-years, respectively. A total 3.8 with 3.8 hypertensive mortality using a random-effects Cox proportional hazards model with 4 degrees of freedom (the minimum AIC value) showed that there was a non-linear association between PM$_{2.5}$ concentration and higher. Specifically, the hypertension risk was significantly greater in individuals with higher concentrations were strongly related with hypertension incidence (AHR: 1.12; 95% CI: 1.09–1.16).

The Cox proportional hazards model with 4 degrees of freedom (the minimum AIC value) showed that there was a non-linear association between PM$_{2.5}$ exposure and hypertension incidence (AHR: 1.01, 95% CI: 1.01–1.02) and per 10 µg/m$^3$ increment in PM$_{2.5}$, the concentration was associated with hypertension incidence (AHR: 1.12; 95% CI: 1.09–1.16). Similar results were observed according to quartiles, where higher concentrations were strongly related with hypertension incidence (Q2 AHR: 1.31, 95% CI: 1.13–1.51; Q3 AHR: 1.35, 95% CI: 1.15–1.60; Q4 AHR: 1.83, 95% CI: 1.53–2.17). Trend analysis showed that there was a linear trend between the PM$_{2.5}$ concentration quartile and hypertension incidence (P trend <0.001).

In subgroup analyses, after adjusting for related confounding variables, the effect estimate for each 10 µg/m$^3$ increment in PM$_{2.5}$ concentration was significantly greater in individuals with without a pension, living in a rural setting, and residing in central/western China (P-value for modification effect <0.05 and significant HR value). The detailed modification effects are shown in Table 4.

DISCUSSION

We conducted a 10-year prospective cohort study to examine the association between PM$_{2.5}$ exposure and the risk of hypertension incidence among Chinese elderly. Our results indicated that long-term exposure to PM$_{2.5}$ was significantly associated with hypertension incidence among elderly individuals aged 65–116 years. We found that per 1 and 10 µg/m$^3$ increase in PM$_{2.5}$, the HR of hypertension incidence increased by 1 and 12% among the elderly, respectively. Similarly, there were strong positive links between Q2, Q3, and Q4 with PM$_{2.5}$ exposure and hypertension incidence (Table 3). Adjusting for potential confounding variables such as age, sex, educational attainment, and pension et al., we found that a 1 µg/m$^3$ increment in PM$_{2.5}$ concentration was associated with hypertension incidence [adjusted HR (AHR): 1.01, 95% CI: 1.01–1.02] and per 10 µg/m$^3$ increment in PM$_{2.5}$, the concentration was associated with hypertension incidence (AHR: 1.12; 95% CI: 1.09–1.16). Similar results were observed according to quartiles, where higher concentrations were strongly related with hypertension incidence (Q2 AHR: 1.31, 95% CI: 1.13–1.51; Q3 AHR: 1.35, 95% CI: 1.15–1.60; Q4 AHR: 1.83, 95% CI: 1.53–2.17). Trend analysis showed that there was a linear trend between the PM$_{2.5}$ concentration quartile and hypertension incidence (P trend <0.001). In subgroup analyses, after adjusting for related confounding variables, the effect estimate for each 10 µg/m$^3$ increment in PM$_{2.5}$ concentration was significantly greater in individuals with without a pension, living in a rural setting, and residing in central/western China (P-value for modification effect <0.05 and significant HR value). The detailed modification effects are shown in Table 4.
incidence. Moreover, we observed that individuals with certain characteristics were more likely to be affected by PM$_{2.5}$ exposure based on subgroup analysis.

We observed that the average exposure from 1-year before the event had the strongest effect on the incidence of hypertension, which was shorter than previous studies on the association between PM$_{2.5}$ exposure and health outcomes have reported. For instance, some studies reported that a 3-year average PM$_{2.5}$ exposure was most closely related to mortality (32, 37). Similarly, other studies have verified that an exposure period of $<2$ years might greatly influence cardiovascular health outcomes (38, 39). Thus, our results indicated that the impact of PM$_{2.5}$ exposure was greater than we expected. The findings of our study suggest that interventions to improve air quality may reduce the incidence of hypertension within a short period.

According to the shape of the dose-response relationship between PM$_{2.5}$ and hypertension incidence, we found that there was no safe threshold for PM$_{2.5}$ exposure (Figure 3). Some studies have detected a threshold for PM$_{2.5}$ exposure and negative health outcomes (22, 40), for example, a J-shaped association existed between PM$_{2.5}$ exposure with a threshold concentration of 33 mg/m$^3$, and function disability among the elderly. Other studies report that there is no safe threshold for PM$_{2.5}$ exposure when considering asthma mortality and all-cause mortality risk (32, 41). Currently, the evidence for safe PM$_{2.5}$ exposure with respect to hypertension incidence is limited. According to our findings, even with exposure to low levels of PM$_{2.5}$, the risk of hypertension should not be overlooked. From a public policy perspective, reducing each one unit concentration of PM$_{2.5}$ may be effective for protecting the elderly. Therefore, stricter national policy actions are recommended to improve air quality.
The adverse effects of PM$_{2.5}$ exposure and increased risk for hypertension found in our study are in line with previous studies conducted among middle-aged and elderly people under 80 years of age (19, 21). For example, Wu et al. used a database including 20,927 middle-aged and older participants and reported that the increase in hypertension incidence risk per unit PM$_{2.5}$ exposure was about 4.8 and 6.3% higher among males and females, respectively (19). Lin et al. enrolled 59,456 adults aged 50 years and older from four cohorts in China and showed that each 10 µg/m$^3$ increase in PM$_{2.5}$ concentration increased the HR of hypertension incidence by 14% (17). In addition, the current study showed that the HRs for Q2, Q3, and Q4 for PM$_{2.5}$ exposure were relatively higher than those in other published studies (12). One explanation for this finding was that these participants in our study were older than other studies (mean age was 87.7 years in our paper). Expert proposes that the ability to adapt to air pollution exposure may be reduced due to aging and health status (22). However, one study that investigated elderly individuals residing in Taiwan found no significant association between PM$_{2.5}$ exposure and hypertension (16). Discrepant results were also observed in a Black Women's Health Study that found PM$_{2.5}$ exposure was not significantly associated with measured hypertension among middle-aged and elderly women (9). In the present study, the mean of PM$_{2.5}$ exposure concentration (55.3 µg/m$^3$) was relatively higher than PM$_{2.5}$ concentration from other studies conducted in the USA (8.6 µg/m$^3$) (7), Seoul (38.9 µg/m$^3$) (42), Europe (range: 5.0–8.9 µg/m$^3$) (43), and a cross-sectional study from India (33.0 µg/m$^3$) (11). Our finding was supported by a prior paper reported that the risk of PM$_{2.5}$ exposure was declining in most western developed countries; in contrast, the risk of PM$_{2.5}$ exposure was prominent in developing countries (44). Since 2013, the Chinese government has carried out a pilot project to monitor PM$_{2.5}$ concentration in 33 major cities and implemented an air quality standard for PM$_{2.5}$ concentration in 2016 (45). In 2019, the PM$_{2.5}$ concentration in 22 of 31 major cities in mainland China exceeded the annual limitation concentration for China (35.0 µg/m$^3$) and 31 of them exceeded the WHO recommendation (10.0 µg/m$^3$) (46). Overall, the regional development patterns and air pollutant concentrations varies across countries and regions, which may explain some of the heterogeneity of PM$_{2.5}$ exposure; the inconsistency between our results and those of other studies may also be caused by differences in time spent outdoors, population structures, and accessibility to health care (32).

To date, relevant studies have been proposed to explain the impacts of PM$_{2.5}$ exposure on blood pressure through several mechanisms. One study posits that PM$_{2.5}$ is taken into the human body by direct translocation through the olfactory bulb, leading to inflammatory responses and oxidative stress (47). Another proposes that inhaled particles destroy the balance in the autonomic nervous system, inducing a sympathetic response, followed by arterial vasoconstriction (48). Finally, it may be that long-term exposure to PM$_{2.5}$ affect the systemic hemodynamics of the body due to endothelial injury or dysfunction, which is also considered a risk factor for hypertension (10). In summary, the beneficial efforts to reduce PM$_{2.5}$ emission on hypertension among the elderly would be important.

Our findings implied that the association between each 10-µg/m$^3$ increment in PM$_{2.5}$ exposure and hypertension was modified by different characteristics of the elderly. We found that HR seemed more apparent in individuals without pensions. One study concluded that the health impact of PM$_{2.5}$ on individuals...
TABLE 3 | The association between long-term exposure to PM2.5 and hypertension incidence among the elderly.

| Models          | PM2.5          | P for trend | PM2.5         |
|-----------------|----------------|-------------|---------------|
|                 | q1  | q2  | q3  | q4  | Per 1-µg/m³ increment | Per 10-µg/m³ increment |
| Crude model     | 1   | 1.31| 1.35| 1.77| <0.001 | 1.01 | 1.12 |
| (1.13–1.51)     |     | (1.15–1.60) | (1.49–2.09) |     | (1.01–1.12) | (1.09–1.16) |
| Adjusted Model 1| 1   | 1.31| 1.35| 1.83| <0.001 | 1.01 | 1.12 |
| (1.13–1.51)     |     | (1.15–1.60) | (1.53–2.17) |     | (1.01–1.12) | (1.09–1.16) |

Crude model: random-effects Cox proportional hazards model for the sampling sites and only included PM2.5 exposure.
Adjusted Model 1: random-effects Cox proportional hazards model for the sampling sites further adjusted for age, gender, education attainment, pension, living arrangement, marital status, regions, residence, smoking at the present, drinking at the present, exercising at the present, BMI index, self-reported diabetes, function disability, GDP per capita, the proportion of secondary industry and heart disease.

TABLE 4 | Subgroup analysis for hypertension incidence and its associated with each 10 µg/m³ increment in PM2.5 concentration.

| Characteristic          | Groups                        | Person year | Cases | Incidence of hypertension (per 100 person year) | HR (95%CI) | P-value for effect modification |
|-------------------------|-------------------------------|-------------|-------|-------------------------------------------------|------------|---------------------------------|
| Age                     | 65–80                         | 9,663.53    | 1,014 | 10.49                                           | 1.26 (1.21–1.32) | (Reference) |
|                         | >80                           | 14,558.56   | 1,052 | 7.23                                            | 0.98 (0.94–1.02) | 0.999 |
| Gender                  | Male                          | 11,096.59   | 937   | 8.44                                            | 1.10 (1.05–1.44) | (Reference) |
|                         | Female                        | 13,125.50   | 1,129 | 8.60                                            | 1.06 (1.02–1.10) | 0.921 |
| Education attainment    | 0                             | 13,968.92   | 1,153 | 8.25                                            | 1.07 (1.03–1.11) | (Reference) |
|                         | >0                            | 10,253.17   | 913   | 8.90                                            | 1.10 (1.05–1.15) | 0.145 |
| Living arrangement      | Living in home with family member(s) | 20,241.03 | 1,703 | 8.41                                            | 1.11 (1.07–1.15) | (Reference) |
|                         | Nursing institution/alone     | 3,981.06    | 363   | 9.12                                            | 1.04 (0.98–1.10) | 0.974 |
| Pension                 | Yes                           | 4,330.60    | 340   | 7.85                                            | 1.00 (0.94–1.07) | (Reference) |
|                         | No                            | 20,241.03   | 1,726 | 8.53                                            | 1.15 (1.11–1.19) | <0.001 |
| Marital status          | Married                       | 9,423.61    | 918   | 9.74                                            | 1.15 (1.10–1.20) | (Reference) |
|                         | Widowed/ Separated/Divorced/ Never married | 14,798.48 | 1,148 | 7.76                                            | 1.03 (0.99–1.07) | 0.999 |
| Regions                 | Eastern                       | 9,560.80    | 819   | 8.57                                            | 1.02 (0.97–1.07) | (Reference) |
|                         | Central/Western               | 14,661.28   | 1,247 | 8.51                                            | 1.22 (1.16–1.27) | <0.001 |
| Residence               | Urban                         | 4,428.99    | 268   | 6.05                                            | 1.01 (0.94–1.08) | (Reference) |
|                         | Rural                         | 19,793.10   | 1,798 | 9.08                                            | 1.16 (1.12–1.20) | 0.001 |
| Smoking at the present  | No                            | 19,439.40   | 1,656 | 8.52                                            | 1.09 (1.05–1.12) | (Reference) |
|                         | Yes                           | 4,782.69    | 410   | 8.57                                            | 1.09 (1.02–1.16) | 0.500 |

(Continued)
TABLE 4 | Continued

| Characteristic                       | Groups   | Person year | Cases | Incidence of hypertension (per 100 person year) | HR (95%CI)       | P-value for effect modification |
|-------------------------------------|----------|-------------|-------|-----------------------------------------------|------------------|---------------------------------|
| Drinking at the present             | No       | 19,417.00   | 1,602 | 8.25                                          | 1.09 (1.05–1.13) | (Reference)                    |
|                                      | Yes      | 4,805.09    | 464   | 9.66                                          | 1.08 (1.02–1.15) | 0.606                           |
| Exercising at the present           | No       | 16,774.29   | 1,389 | 8.28                                          | 1.10 (1.06–1.14) | (Reference)                    |
|                                      | Yes      | 7,447.80    | 677   | 9.09                                          | 1.07 (1.02–1.12) | 0.864                           |
| Self-reported diabetes               | No       | 23,806.04   | 2,013 | 8.46                                          | 1.12 (1.09–1.16) | (Reference)                    |
|                                      | Yes      | 416.05      | 35    | 8.41                                          | 1.20 (0.90–1.59) | 0.337                           |
| Self-reported heart disease          | No       | 22,814.56   | 1,958 | 8.58                                          | 1.12 (1.08–1.16) | (Reference)                    |
|                                      | Yes      | 1,407.53    | 108   | 7.67                                          | 1.10 (0.99–1.23) | 0.644                           |
| BMI index                            | 18.5–23.9 | 13,282.30   | 1,143 | 8.61                                          | 1.09 (1.05–1.14) | (Reference)                    |
|                                      | <18.5    | 8,242.74    | 647   | 7.85                                          | 1.03 (0.97–1.08) | 0.974                           |
|                                      | >23.9    | 2,697.05    | 276   | 10.23                                         | 1.07 (1.00–1.15) | 0.761                           |
| Function disability                  | Yes      | 3,341.052   | 137   | 4.10                                          | 1.04 (0.95–1.14) | (Reference)                    |
|                                      | No       | 20,881.036  | 1,929 | 9.24                                          | 1.13 (1.09–1.17) | 0.955                           |
| GDP per capita                       | <P_{50}  | 12,129.68   | 1,085 | 8.78                                          | 1.30 (1.23–1.38) | (Reference)                    |
|                                      | ≥P_{50}  | 12,092.41   | 1,001 | 8.28                                          | 1.05 (1.01–1.10) | 0.999                           |
| The proportion of secondary industry | <P_{50}  | 11,237.03   | 988   | 8.79                                          | 1.25 (1.18–1.32) | (Reference)                    |
|                                      | ≥P_{50}  | 12,985.06   | 1,078 | 8.30                                          | 1.08 (1.04–1.13) | 0.999                           |

The bold values mean statistic significance such as P = 0.001.

with lower socioeconomic status was slightly greater than that of others (49). In our study, most of the participants were uneducated and had no pension, so they were probably more likely to spend their time engaging in outdoor manual labor that more easily suffered from PM$_{2.5}$ exposure. Additionally, the subgroup analysis showed that respondents who lived in rural areas and central/western China were more likely to have hypertension. According to the distribution of PM$_{2.5}$ concentration in mainland China in this study (Figure 2), we found most of the central or western areas such as Szechwan, Hunan, and Hubei provinces, and the Beijing-Tianjin-Hebei region had higher air pollutant levels. It is important to note that these areas are traditional industrial districts with a high population density. In rural areas, the PM$_{2.5}$ concentration was higher than in urban areas probably due to residential energy materials such as coal and biomass fuels. Accordingly, several approaches, such as clean energy use in rural areas and upgrading of traditional industries in traditional industrial districts, would improve air quality and decrease the risk of hypertension among the elderly.

Some limitations of this study should be acknowledged. First, according to the diagnostic guide, antihypertensive medication was one of the criteria for hypertension diagnosis, but the information was not available in the CLHLS database. In our study, we used blood pressure values and self-reported history of hypertension diagnosis to identify hypertension, which might be affected by recall bias. However, it could be confirmed that individuals diagnosed with hypertension by a physician should be prescribed antihypertensive medication (16). The address of each participant was deleted from the CLHLS database and we could not collect the precise addresses of participants from this publicly open database. However, our study estimated PM$_{2.5}$ exposure in residential counties or district addresses using satellite observations and an atmospheric chemistry database.
that effectively employed relatively coarse spatial resolution to estimate large-area variations in air pollution exposures (11, 50). Besides, lots of factors such as economic status, disease comorbidities, and lifestyle, which were dynamic and not easy to assess. Lastly, because the concentrations of other air pollutants such as \( \text{SO}_x \), \( \text{NO}_x \), and \( \text{O}_3 \) were not unavailable in the open database, we could not confirm the effects of PM\(_{2.5}\) exposure and hypertension incidence under the influence of these air pollutants. A more comprehensive evaluation should be conducted to verify the influence of various air pollutants on hypertension in future studies.

Despite these limitations, this study further added to the present evidence of the effects of PM\(_{2.5}\) exposure on hypertension among the elderly. In particular, the study provided quantitative evidence relying on a 10-year period using a large population-based cohort that included a nationwide representative sample of a specific age group (\( \geq 80 \) years) in a developing country. The results can be used to inform public policy to address the urgent need for more rigorous environment-related policies for air pollution abatement and make efforts to protect vulnerable elderly away from such harmful exposure.

**CONCLUSION**

This study highlighted the effects of long-term exposure to PM\(_{2.5}\), which was associated with hypertension incidence among the elderly, and indicated that there was no safe threshold for the association between PM\(_{2.5}\) and risk of hypertension. Furthermore, this relationship was greater in the elderly who were without pension, living in rural areas, and residing in central/western China. Reduction approaches and policies for PM\(_{2.5}\) should be developed to reduce the incidence of hypertension and improve the quality of life of the elderly.

**DATA AVAILABILITY STATEMENT**

The CLHLS questionnaires are available at https://sites.duke.edu/centerforaging/programs/chinese-longitudinal-healthy-longevity-survey-clhls/. The full datasets used in this analysis are available from the corresponding author upon reasonable request.

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**ETHICS STATEMENT**

The studies involving human participants were reviewed and approved by the Research Ethics Committees of Duke University and Peking University approved the protocol for each wave of the CLHLS, which is the data source of this study. The survey respondents gave informed consent before participating. The patients/participants provided their written informed consent to participate in this study.

**AUTHOR CONTRIBUTIONS**

ZW and ZF: data cleaning. ZW, ZF, CW, and LL: conceptualization and visualization. ZW, CW, ZF, and WW: data curation, writing, original draft preparation, methodology, software, and reviewing. ZW, ZF, CW, LL, and WW: supervision. All authors contributed to the article and approved the submitted version.

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**SUPPLEMENTARY MATERIAL**

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