New type II Cepheids from VVV data towards the Galactic center

V. F. Braga\textsuperscript{1,2}, R. Contreras Ramos\textsuperscript{1,3}, D. Minniti\textsuperscript{1,2,4}, C. E. Ferreira Lopes\textsuperscript{5}, M. Catelan\textsuperscript{1,3,\textsuperscript{**}}, J. H. Minniti\textsuperscript{1,3,6}, F. Nikzad\textsuperscript{1,3}, and M. Zoccali\textsuperscript{1,3}

\textsuperscript{1} Instituto Milenio de Astrofísica, Santiago, Chile
\textsuperscript{2} Departamento de Física, Facultad de Ciencias Exactas, Universidad Andrés Bello, Fernández Concha 700, Las Condes, Santiago, Chile
\textsuperscript{3} Pontificia Universidad Católica de Chile, Instituto de Astrofísica, Av. Vicuña Mackenna 4860, 7820436 Macul, Santiago, Chile
\textsuperscript{4} Vatican Observatory, V00120 Vatican City State, Italy
\textsuperscript{5} National Institute For Space Research (INPE/MCTI), Av. dos Astronautas 1758, São José dos Campos, SP 12227-010, Brazil
\textsuperscript{6} European Southern Observatory, Alonso de Córdova 3107, Casilla 19001, Santiago, Chile

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ABSTRACT

Context. The Galactic center (GC) is the densest region of the Milky Way. Variability surveys towards the GC potentially provide the largest number of variable stars per square degree within the Galaxy. However, high stellar density is also a drawback due to blending. Moreover, the GC is affected by extreme reddening, therefore near infrared observations are needed.

Aims. We plan to detect new variable stars towards the GC, focusing on type II Cepheids (T2Cs) which have the advantage of being brighter than RR Lyrae stars.

Methods. We perform parallel Lomb-Scargle and Generalized Lomb-Scargle periodogram analysis of the \(K_s\)-band time series of the VISTA variables in the Vía Láctea survey, to detect periodicities. We employ statistical parameters to clean our sample. We take account of periods, light amplitudes, distances, and proper motions to provide a classification of the candidate variables.

Results. We detected 1019 periodic variable stars, of which 164 are T2Cs, 210 are Miras and 3 are classical Cepheids. We also found the first anomalous Cepheid in this region. We compare their photometric properties with overlapping catalogs and discuss their properties on the color-magnitude and Bailey diagrams.

Conclusions. We present the most extensive catalog of T2Cs in the GC region to date. Offsets in \(E(J - K_s)\) and in the reddening law cause very large (~1–2 kpc) uncertainties on distances in this region. We provide a catalog which will be the starting point for future spectroscopic surveys in the innermost regions of the Galaxy.

Key words. stars: variables: Cepheids – Galaxy: bulge – Galaxy: center

1. Introduction

To investigate the central regions of the Galaxy is a top priority for studies of Galaxy evolution, since the Bulge contains \(\sim40\%\) of the stellar mass of the Galaxy (Valenti et al. 2016; McMillan 2017). However, the Galactic center (GC) is one of the most complicated regions to study in the Milky Way because of high extinction and high crowding. At low galactic latitudes, the optical surveys such as \textit{Gaia} (Gaia Collaboration 2016), OGLE (Udalski et al. 1992), and ASAS (Pojmanski 1997) suffer from extreme reddening and need to be complemented with near infrared (NIR) surveys like the Two Micron All-Sky Survey (2MASS; Skrutskie et al. 2006) and the VISTA Variables in the Vía Láctea (VVV; Minniti et al. 2010; Saito et al. 2012) to detect sources in the regions of the Galactic plane where visual absorption can be as high as \(A_V \sim 30\) mag, or higher.

Over the last 10 years the synergy between OGLE and VVV has provided unprecedented results concerning variable stars in the Bulge thanks to the exploitation of their advantages and compensation for their disadvantages. In fact, OGLE scanned vast regions of the sky, generating time series with thousands of epochs, and providing the largest known catalogs of variable stars in the Bulge and Magellanic Clouds. However, the surface density of both periodic variables (RR Lyraes [RRLs], Cepheids, Miras...) and transients (microlensing events) drops dramatically at low galactic latitudes (\(|b| < 1.5\,^\circ\)). On the other hand, VVV could not achieve the same sky area coverage and number of epochs due to the intrinsically more telescope time-demanding observation strategies in the NIR bands. Nonetheless, its deeper images allowed the detection of Cepheids, RRLs, and microlensing events in regions where none had been found before (Dékány et al. 2013; Gran et al. 2016; Minniti et al. 2016, 2017a; Navarro et al. 2017, 2018; Majaess et al. 2018; Contreras Ramos et al. 2018, henceforth, CR18), also leading to the discovery of new globular clusters (Minniti et al. 2017bc; Bica et al. 2018). Concerning pulsating variable stars in the instability strip (IS), which are valuable standard candles since they obey period–luminosity (PL) relations, the synergy between OGLE and VVV was discussed by Pietrukowicz et al. (2012) and exploited by Bhardwaj et al. (2017a) and Braga et al. (2018) for type II Cepheids (T2Cs). However, since these studies were based on the OGLE lists of variables, the inner regions of the Bulge could not be inspected.

In this framework, a census of T2Cs at low latitudes is still missing. These variables are not as popular as classical Cepheids...
(CCs) or RRLs as standard candles. In fact, they are ∼2–3 mag fainter than CCs at the same period – meaning that they cannot be used to estimate distances of stellar systems outside the Local Group – but they are also much less numerous than RRLs (less than 1000 were found in the Bulge by OGLE, while the same survey detected more than 38,000 RRLs Soszyński et al. 2014, 2017). On the other hand, T2Cs are brighter than RRLs. Moreover, the majority of the T2Cs are old (>10 Gyr) stars, although it was recently suggested that the W Vir (WV) subclass might partly be associated to the intermediate-age population (Iwanek et al. 2018). This means that they can also be found in stellar systems with no recent star formation events, which is a requirement to find CCs as they are purely young (<400 Myr, Bono et al. 2005; Anderson et al. 2016) stars. For a recent review of the properties of T2Cs, see the monography by Catelan & Smith (2015).

The aim of this paper is to find new variables in the centralmost tile of the VVV survey, focusing on the detection of T2Cs. The paper is organized as follows: in Sect. 2, we present the data. In Sect. 3, we describe the light-curve analysis, including the variable star search, period determinations, and light-curve fittings. In Sect. 4, we discuss the classification of periodic variables using all tools available. We present the matches with other existing catalogs of GC variable stars in Sect. 5. In Sect. 6, we discuss the position, color-magnitude diagrams (CMDs) and Bailey (amplitude vs. period) diagrams of the final list of variables, and the obtained distances to T2Cs. Finally, the conclusions are presented in Sect. 7.

2. Data

Photometry. We used proprietary PSF photometric reduction (Contreras Ramos et al. 2017) of the VVV data (Minniti et al. 2010) in tile b333 (~1:23 < l < 0:23; ~0:46 < b < 0:65), which covers 1.501 deg² on the GC. The camera VIRCAM has 16 chips which cover one full tile with six pointings, which are referred to as “pawprints”. Overall, the observation of each tile consists of 96 fields of view. We refer to each of these fields of view as a “pawprint-chip combination” (PCC). The ID of each PCC is a three-digit number, where the first digit is the pawprint number, and the second and third are the chip number. We note that since there are overlaps between the different PCCs, one tile is fully covered by 48 PCCs. The advantages of PSF-fitting photometry with respect to the public aperture photometry available at VSA¹ are manifold in crowded regions like b333, which includes the GC; it allows for blended sources to be resolved more easily, performs a better sky subtraction, and provides a deeper limit magnitude (at least ~2 mag, and up to ~3 mag in the most crowded regions, like b333). In principle, PSF photometry is less accurate than aperture photometry for bright, saturated sources. However, this effect is mitigated in a field like b333 for two reasons: First, because of the extreme reddening, the bright and saturated sources are statistically less frequent. Second, the aperture photometry of bright sources is performed by adopting large apertures, but the extreme crowding of b333 would hamper the results provided by this method.

Our photometric data set consists of 104 epochs in the $K_s$ band and two epochs in each of the $ZYJH$ bands. We consider the full $K_s$-band time series and the average magnitudes in the other filters.

Reddening. We adopted the 3D reddening map of Schultheis et al. (2014; henceforth S14). This map provides a 6 × 6 arcmin grid of $E(J - K)$ values in 21 bins of distance, from 0 to 10.5 kpc. The reddening provided by this map depends not only on the coordinates but also the distance of the target. Moreover we could not estimate the distance of all targets because not all of them are reliable distance indicators (e.g., binaries, RV Tau (RVT) stars, non-pulsating variable stars). For these variables, we have adopted the 2D map by Gonzalez et al. (2012; henceforth G12), which is independent of distance. The $E(J - K_s)$ values provided by this map are optimal for targets at the distance of the GC (~8.3 ± 0.2[stat.] ± 0.4[syst.] kpc, de Grijs & Bono 2016) but are overestimated for targets closer than the GC, and vice versa (Schultheis et al. 2014; Braga et al. 2018a). To estimate the extinction $A_{ZYJH}$ in the single bands, we adopted the reddening law of Alonso-Garcia et al. (2017). We note that the G12 and S14 reddening maps and the adopted reddening law were obtained using VVV data. Finally, let us mention that for T2Cs, CCs, and anomalous Cepheids (ACs), we also derive independent reddening estimates by simultaneously solving the PLH and PLKs relations for distance and reddening.

Proper motions. We adopt proper motions ($μ_l$, $μ_b$) from the PSF photometry itself and $μ_l$ and $μ_b$ were estimated with the same method used in Contreras Ramos et al. (2017). We note that these proper motions are relative to the Galaxy. We tried to match the targets in our final list of candidate variable stars with the Gaia DR2 (Gaia Collaboration 2016, 2018) source catalog, but we found reliable matches for only 45 stars, of which only 30 have a five-parameter astrometric solution. Since these represent the minority of our sample of candidate periodic variable stars (PVSs) and parallaxes are negative for almost half of the sample, indicating an uncertain astrometric solution, we did not use these data. This was expected since Gaia performances in extremely crowded regions are not optimal.

3. Analysis of the light curves

We found 5 147 696 sources with $K_s$-band light curves within the tile. Such an amount of data requires proper selection to be manageable, and therefore we adopted rejection criteria at several steps of the data analysis.

3.1. Preliminary target selection

As a first step, we rejected all the sources that have light curves with either few phase points or poor-quality photometric reduction, as well as those that are considered to be nonvariable within the photometric uncertainties. More specifically, we rejected those for which (a) the light curve has less than 25 phase points with a valid PSF reduction output (Gran et al. 2015; Molina et al. 2019); (b) the $χ^2$ of the PSF fitting (Stetson 1994) is larger than 3.5, $χ^2$ which indicates the difference between the pixel counts in the CCD and the PSF that fits the source. We noted that the sources with $χ^2 > 3.5$ trace a “plume” in the sharpness ($sha$)-magnitude plane (see Fig. 1). The $sha$ parameter indicates whether the profile of the source is broader or narrower than the PSF profile, and is useful to detect nonstellar sources, which have large absolute values of $sha$; (c) the difference between the maximum and the minimum magnitude of the time series $ΔK_s$ is smaller than 0.1 mag. We note that this is not a cut in the $K_s$-band amplitude (Amp($K_s$)). In fact, due to the scatter of the phase points, $ΔK_s$ is always larger than Amp($K_s$) and the quoted cut still allows us to detect variables with Amp($K_s$) down to ~0.05 mag; (d) the median uncertainty on the phase points is larger than 0.2 mag. We note that a magnitude uncertainty of 0.2 mag means

¹ http://horus.roe.ac.uk/vsa/
regions of the plane) to red (densest). The thick dashed line in the
Fig. 1.

ED magnitudes are taken into account, and not the epochs of obser-
tive of the light curve of a PVS. We note that only the measured
must be an array of observed magnitudes – could be representa-
shape parameters for distributions with an even number of items
are more precise than 20% (point d). In the following, we discuss
variability over noise (point c), and measured magnitudes which
tudes (point b), an amplitude which is large enough to detect
photometric solution for the accuracy of the measured magni-
tation in the LS is fixed at the empirical mean magnitude of the phase
of each of the seven EVPs, the calculated \( I_{p,i} \), as defined in
Ferreira Lopes & Cross (2017; Eq. (19)). By construction, \( I_{p,i} < 1 \)
indicates a nonvariable star according to the ith EVP. One may
also take account of the seven EVPs altogether and reject – as nonPVSs – all the targets for which the sum of the \( I_{p,i} (\Sigma_{p}) \) is
less than seven. We note that this is possible because for a
given target the seven \( I_{p,i} \) are all of the same order of magni-
However, we have empirically checked that there are some
fide PVSs – especially those with small amplitudes – for
which \( \Sigma_{p} < 7 \) (see Fig. 2 which shows that several candidate
PVSs are below the \( I_{p,i} = 1 \) locus.). Therefore, we decided to
adopt a more conservative approach and we kept, as candidate
PVSs, all the targets for which \( \Sigma_{p} > 3.3 \). This threshold was
chosen as a compromise between a moderate increase of false
alarms and a higher completeness.

3.2. Preliminary phase point selection

After the rejection based on the criteria (a), (b), (c), (d) and (e)
in Sect. 3.1, where ~15% of the sources were discarded, for
each light curve we performed a rejection of the phase points
that we assumed to be outliers. The rejection process was based
on an iterative sigma clipping at a 3.5\( \sigma \) level with respect to the
median magnitude. This is a conservative threshold that allowed
us to detect periodic variable stars with large amplitudes (up to
~2 mag) and at the same time to discard outlying phase points
that could affect the accuracy of the periodicity search. In
this step, for ~57% of the light curves, we do not reject any phase
points. For ~23%, ~10%, ~5%, and ~2% of the light curves,
the number of points that were rejected was one, two, three,
and four, respectively. For the remaining ~3% of the light curves,
five or more points (up to eighteen) were rejected.

3.3. Periodicity search

After the selections and cuts described in the above section, we
ran our periodicity search algorithm over all the light curves.
As a preliminary step, we converted our epochs from Julian dates
into heliocentric Julian dates (\( \eta_{HJD} \)). We calculated both the
classical Lomb-Scargle (LS; Scargle 1982) and the Generalized Lomb-
Scargle (GLS; Zechmeister & Kürster 2009) periodograms. The
difference between the two is that the zero point of the model func-
tion in the LS is fixed at the empirical mean magnitude of the phase
points, while in the GLS method, the zero point is allowed to float.
Therefore, the GLS is more reliable when the light curve is not
well sampled (VanderPlas & Ivezic 2015).

We performed both a LS and a GLS analysis because the
comparison of the periodograms is useful to detect and automatic-
ly reject stars that are not periodic variables. We have veri-
fied empirically that, if the highest peak of the GLS periodogram
does not match – within the typical peak width – any among the
ten highest peaks of the LS periodogram, the peaks are asso-
ciated with light variations that are not strictly periodical, and
therefore we can reject the target as a nonPVS. In the following
sections, we explain the details of our analysis.

Fig. 1. \( \chi^2 \) (top panel) and \( \text{sh}a \) (bottom panel) parameters as a function of
K\(_v\)-band magnitude. The density map goes from blue (least dense
regions of the plane) to red (densest). The thick dashed line in the top
panel shows the \( \chi^2 \) cut applied, to reject all sources for which the photometric
solution is poor. These sources are marked as black points in the
bottom panel.

a ~20% uncertainty on the flux; and (e) the variability indices
based on even statistics (Ferreira Lopes & Cross 2016, 2017)
indicate a nonvariable nature for the star. We did not reject satu-
rated (\( K_v < 12 \text{ mag} \), Contreras Ramos et al. 2017) stars because,
although mean magnitudes and amplitudes might be inaccurate,
periods are unaffected by saturation, at least for variables with
Amp(\( K_v \)) \( \gtrsim 0.2 \text{ mag} \). Moreover, we are interested in matching
our variable stars with Miras, CCSs, and T2Cs found by Matsunaga
et al. (2009, 2013). Their survey was shallower than the VVV due
to the smaller diameter of their telescope (IRSF, 1.4 m); there-
therefore all their variables have mean \( K_v \) magnitudes brighter than
14 mag and more than half are brighter than 12.5 mag. We
discuss the offset in mean magnitude due to saturation in more detail
in Sect. 5.5.

The quoted criteria are driven by the requirements to have
enough points to obtain a solid periodogram (point a), a good
photometric solution for the accuracy of the measured magni-
tudes (point b), an amplitude which is large enough to detect
variability over noise (point c), and measured magnitudes which
are more precise than 20% (point d). In the following, we discuss
point (e) in more detail.

Ferreira Lopes & Cross (2017) presented new dispersion
and shape parameters for distributions with an even number of items
in order to decipher whether or not the input distribution – which
must be an array of observed magnitudes – could be representa-
tive of the light curve of a PVS. We note that only the measured
magnitudes are taken into account, and not the epochs of observation.
We have adopted the seven parameters \( ED_{cry}, ED_{era}, ED_{y}, ED_{ma}, ED, ED_{11}, \) and \( ED_{21} \), as defined in Ferreira Lopes & Cross (2017; Table 1).
To avoid ambiguity, hereafter we refer to these seven parameters as EVPs, where \( r \) ranges from 1 to 7.

As a first step, for each of the seven EVP, in each of
the 48 PCCs, we derived the modified Strateva noise models
(Ferreira Lopes & Cross 2017, Eq. (18)) as a function of the
mean magnitude. We adopted the modified models instead of the
classical ones (Strateva et al. 2001; Sesar et al. 2007) because
the former better follow the distribution in the EVP-versus-
magnitude plane, especially at bright magnitudes (see Fig. 4 in
Ferreira Lopes & Cross 2017). We note that the modified Strat-
eva models that we derived also correctly follow the distribu-
tion at faint magnitudes (see Fig. 2, although Ferreira Lopes &
Cross (2017) pointed out that the modified Strateva models at
faint magnitudes deviate from the distribution.

For each of the EVPs, we calculated \( I_{p,i} \), as defined in
Ferreira Lopes & Cross (2017; Eq. (19)). By construction, \( I_{p,i} < 1 \)
indicates a nonvariable star according to the ith EVP. One may
also take account of the seven EVPs altogether and reject – as nonPVSs – all the targets for which the sum of the \( I_{p,i} (\Sigma_{p}) \) is
less than seven. We note that this is possible because for a
given target the seven \( I_{p,i} \) are all of the same order of magni-
However, we have empirically checked that there are some
fide PVSs – especially those with small amplitudes – for
which \( \Sigma_{p} < 7 \) (see Fig. 2 which shows that several candidate
PVSs are below the \( I_{p,i} = 1 \) locus.). Therefore, we decided to
adopt a more conservative approach and we kept, as candidate
PVSs, all the targets for which \( \Sigma_{p} > 3.3 \). This threshold was
chosen as a compromise between a moderate increase of false
alarms and a higher completeness.
To calculate the periodograms, we used an evenly spaced grid of frequencies, from $v_{\text{min}} = \frac{1}{\Delta t}$ days$^{-1}$ – where $\Delta t$ is the total time between the first and the last epoch of the light curve in days – to $v_{\text{max}} = 0.995$ days$^{-1}$. This range was chosen based on the following considerations. (1) Our first aim is to detect T2Cs and fundamental-mode (FU) CCs, which have periods from $\sim$1 to $\sim$200 days. (2) We did not include frequencies that are too close to 1 day$^{-1}$ because the increase in the number of genuine variables would come at the cost of a much higher increase of aliases around 1 day. Based on the typical number of phase points for our light curves, the limit at 0.995 days$^{-1}$ allows us to automatically reject most of the aliases. (3) We did not extend $v_{\text{max}}$ to the range of frequencies of RRLs because the computation time would be around ten times longer with $v_{\text{max}} = 5$ days$^{-1}$. Although the limit at $v_{\text{max}}$ might generate aliases, we adopted an a posteriori LS analysis on a much smaller sample to solve the aliases (see Sect. 3.4.3).

The number of points in the frequency grid of the periodogram ($N_{\text{freq}}$) depends on the range of epochs covered by the light curve ($\Delta t$). As a matter of fact, $DF = \Delta t^{-1}$ is the characteristic width of a peak in the periodogram. Therefore, we set $N_{\text{freq}} = \text{int}(n_{\text{cpp}} \cdot \Delta t \cdot (v_{\text{max}} - v_{\text{min}}))$, where $n_{\text{cpp}}$ is the oversampling factor (Graham et al. 2013) and $n_{\text{cpp}}$ corresponds to a phase shift of about 0.1 between neighboring frequencies that ensure the signal detection of almost all variable types (for more detail see Ferreira Lopes et al. 2018). We set $n_{\text{cpp}} = 9$ for all the light curves (normally, the recommended value of $n_{\text{cpp}}$ is between 5 and 10; VanderPlas & Ivezić 2015). Since $\Delta t$ is not the same for all the light curves, $N_{\text{freq}}$ ranges between $\sim$3000 and $\sim$16 600, with a mean of 16 082.

### 3.3.2. Height of the peaks and frequency aliases

After obtaining the periodograms, we empirically verified that if the peaks are not high enough or if there are too many low peaks – see details below – then the light curve is too noisy to properly classify our targets.

Therefore, we perform a further selection of targets based on the intensity of the peaks of the periodograms. Adopting the standard deviation of the periodogram itself ($\sigma_P$), we discarded (1) light curves with less than 30 points and no peaks higher than $8\sigma_P$; (2) light curves with 30 points or more, and no peaks higher than $9\sigma_P$; and (3) light curves with more than 140 peaks higher than $5\sigma_P$ and no peaks higher than $12\sigma_P$.

After this further selection, we adopted a procedure to check – and eventually correct – aliased frequencies. Following VanderPlas & Ivezić (2015) and VanderPlas (2018), we have checked whether the frequency of the highest peak ($v_{\text{hi}}$) is either a harmonic of the fundamental frequency, an alias, or the true fundamental frequency. This is done in two steps. First, we checked whether or not there were sufficiently high peaks (larger than $6\sigma_P$ or $7\sigma_P$ if the number of phase points is, respectively, smaller or larger than 30) around the critical frequencies ($\nu_{\text{check}}$). To check for harmonics, $\nu_{\text{check}} = \frac{\nu_{\text{hi}}}{m}$, with $m = 2$, 3; to check for aliases,

\[ \nu_{\text{check}} = |\nu_{\text{hi}} + n \cdot \nu_{\text{val}}|, \]

where $\nu_{\text{val}} = 1$ day$^{-1}$ and $n = \pm 1, \pm 2$. Second, we compared the $\chi^2$ of the Fourier fits at $\nu_{\text{val}}$ and $\nu_{\text{check}}$, and selected the best frequency as that which minimizes the $\chi^2$.

Finally, we obtained best frequency estimates for 889 663 targets, of which 435 126 and 454 537 are from GLS and LS periodograms, respectively.

We folded the light curves calculating the phases $\phi$ as the decimal part of $\frac{HJD - T_0}{P_{\text{best}}}$, where $T_0$ is an arbitrary zero epoch that we set at 0.0 days.

### 3.4. Light-curve fit

Despite the quoted cuts and selections, after the analysis of the periodograms, we still have almost one fifth from the original
After obtaining the Fourier fit of the light curves, we derived and Fourier series of higher order would provide unreliable fits variable stars. We use the second order because, at this stage, the targets (mean magnitude, amplitude, and uncertainties), we fitted around the maximum and the minimum (Braga et al. 2018b). The points around the maximum and minimum, plus the standard deviation of the phase points around the fit plus the median photometric error on the phase points.

The uncertainty on the amplitude was derived as the sum in quadrature of the standard deviation of the phase points around the maximum and the minimum, plus the standard deviation of these phase points around the fit of the light curve.

The final value was weighted with the number of phase points around the maximum and the minimum (Braga et al. 2018b). The fits might show, at most, two secondary minima and maxima. In this case, we estimate the amplitude of the bump \( \text{Amp}(K_{\text{bump}}) \) as the difference between the secondary minimum and maximum.

### 3.4.2. Selection of the final sample

With almost 900,000 targets remaining, we applied further selections aimed to narrow down the sample to a more manageable size.

**Selection on phase gaps.** A fraction of folded light curves display wide gaps in phase (≥0.25 cycles). We empirically checked that for stars with \( P_{\text{test}} \leq 330 \) days, these gaps are caused by a poor estimate of \( P_{\text{test}} \) and that the target is not a PVS. Therefore, we reject all targets with a phase gap wider than 0.25 cycles and \( P_{\text{test}} \leq 330 \) days. Targets with wide gaps but longer periods are kept in the sample of candidate PVSs because bona fide PVSs with long periods also show such gaps, caused by the 1 year alias.

**Selection on amplitude.** We put an upper limit of \( \text{Amp}(K_{\text{f}}) \leq 3.5 \) mag on the light amplitude. This is a very conservative upper limit for the light amplitudes of Miras, which are the variables with the largest amplitudes among those that we are interested in. This threshold is based on photometric surveys of Miras in the Bulge, both in the NIR (Matsunaga et al. 2009, which does not find any Mira with \( \text{Amp}(K_{\text{f}}) \geq 2.7 \) mag), and in the optical (Sozzi et al. 2013, which does not find any Mira with \( \text{Amp}(I) \geq 7 \) mag; we note that for Miras we can assume a ratio \( \text{Amp}(K_{\text{f}})/\text{Amp}(I) \approx 0.45 \) (Whitelock 2012). Moreover, the accuracy and precision of our data do not allow to detect PVSs with \( \text{Amp}(K_{\text{f}}) < 0.03 \) mag (Gran et al. 2015). Therefore, we only keep the targets with \( 0.03 \leq \text{Amp}(K_{\text{f}}) \leq 3.5 \) mag.

**Selection on bumps.** Second-order Fourier fits might show local minima and maxima. Although some variables of the IS do show these features, they are not particularly deep, especially at long wavelengths (Lane & Stobie 1993). Therefore, if \( \text{Amp}(K_{\text{bump}}) > \text{Amp}(K_{\text{f}}) \) for a deep bump, we discard the target because these features in the light curve fit show up in noisy or nonperiodic light curves.

**Selection on template fit.** Templates of NIR \((JHK_s)\) light curves of T2Cs are available (Bhardwaj et al. 2017b). For all targets with 1 day \( \leq P_{\text{test}} \leq 80 \) days, we derived the template fit of the light curve. We note that we did not properly apply the template procedure to find the mean magnitude because this would require the knowledge of the epoch of maximum light. We performed a least-squares minimization of the template function \( T(\phi) \) with respect to three independent variables, which are the shift in magnitude (\( \Delta \text{mag} \)), the shift in phase (\( \Delta \phi \)), and the amplitude of the template fit \( AK_{\text{fit}}(\phi) \). We note that in this process the shape of the template fit is fixed. Therefore, the least-squares minimization process only shifts the fitting function horizontally, vertically, or stretches its amplitude. We subsequently compare the template fit and the Fourier fit. If they are similar, this means that the target can be considered as a candidate T2C. Therefore, we compute a parameter \( c_{10} \) to quantitatively compare the two fits, defined as the average of the absolute value of the difference between the two fits, divided by the average amplitude of the two fits: \( c_{10} = \frac{\langle |F(\phi) - T(\phi)| \rangle}{(\text{Amp}(K_{\text{fit}}))} \). If \( c_{10} > 0.2 \), the two fits are quite different from one another, therefore we can discard the target from the list of candidate PVSs. We note that 0.2 is a conservative threshold that allows for targets that are plainly not PVSs to be rejected, but also allows us to keep eclipsing binaries (EBs) with light curves that are quite different from those of T2Cs. Taking account of the accuracy of our data, we can assume that the shape of the light curves of CCs is similar to those of T2Cs. We tried to fit the light curves using also the templates of CCs (Inno et al. 2015), trying to separate between T2Cs and CCs, based on the \( \chi^2 \) of the two fits. However, this was not possible because the two templates are very similar over the whole range of periods. As a matter of fact, for the majority of our light curves (all those with \( K_{\text{f}} \geq 14.5 \) mag) the \( \chi^2 \) of the two fits are often similar. Figure 3 shows the light curve of a variable which is clearly a T2C (based on both its location in the Bailey diagram and its proper motion) despite the \( \chi^2 \) of the T2C template fit being larger than that of the CC template fit (see labels).

**Selection and ranking on modified \( \chi^2 \).** A further selection is based on the dispersion of the phase points around the fit. However, we did not compute a simple \( \chi^2 = \frac{\sum_{i=1}^{n} (\text{mag}_i - F(\phi_i))^2}{n} \), where mag. is the measured magnitude of the \( i \)th phase point and...
Fig. 4. Top: light curve of the variable b333_612_30569. The name and period (in days) are labeled at the top. The Fourier fit is displayed as a magenta line. The χ², Amp(K₂) and χ² s are labeled at the top. Bottom: as in the top panel, but for b333_612_81923.

\( F(\phi_i) \) is the value of the Fourier fit at the phase \( \phi_i \) of the \( i \)th phase point. We adopted a modified chi squared \( \chi^2 = \chi^2 / \text{Amp}(K_2)^2 \) that we also used to rank our final sample of targets for the visual inspection, from the “clearest” (smaller \( \chi^2 \)) to the “least clear light” curves. The division by the squared amplitude has the effect of increasing the \( \chi^2 \) of the small-amplitude variables and in turn decreasing their rank. This operation is needed to obtain a more reliable ranking: Fig. 4 shows two variables that have an identical \( \chi^2 \) but different \( \text{Amp}(K_2) \). Based on \( \chi^2 s \), the variable in the upper panel is ranked higher than the one in the bottom panel to take account of the relative dispersion with respect to the amplitude. Finally, we rejected all the targets for which \( \chi^2 s > 0.2 \).

Visual inspection. After the quoted steps, we were left with 54 667 targets, which is a little more than 1% of the starting sample. This is still too large and contains many stars that cannot be considered as bona fide candidate PVSs. We performed a visual inspection and found 1013 “top tier” candidates. We note that the final catalog includes 1019 targets because we manually added six Cepheids already found in the literature (Matsunaga et al. 2011, 2013) which are present in our initial target list but were not retrieved by the periodicity search algorithm (see Sects. 4 and 5.3).

3.4.3. Refining the periods and fits

After the quoted selections it is possible to visually inspect the light curves of all the candidate variables to improve our analysis. As a matter of fact, these steps are mandatory for a more accurate classification of the PVS candidates.

Unresolved aliases. The alias check process performed in Sect. 3.3.2 has one disadvantage: it does not detect aliases at frequencies outside the frequency grid (in our case, for \( v > 0.995 \) days⁻¹). Therefore, we repeated the whole periodicity search process with LS and GLS on the 1019 PVS candidates, but this time on a frequency grid going from \( v_{\text{min}} = 0.000 \) days⁻¹ to \( v_{\text{max}} = 5 \) days⁻¹. Since the size of this sample is much smaller than the starting list of sources, this is not a time-consuming process. For 77 targets, we found a \( P_{\text{best}} \) shorter than 1 day, which is different from that obtained in the first iteration.

Order of the Fourier fit. Some variables – especially the brightest ones, which have small photometric errors – show well-defined light curves with low noise, and are better fitted by Fourier series of an order higher than the second. For these variables, we repeated the fit using a third-, fourth-, or fifth-order Fourier series. We also repeated the fits of almost sinusoidal light curves using a first-order Fourier fit to avoid unphysical bumps appearing especially in light curves with few phase points.

Frequency doubling. EBs, RVTs, but also WVs with periods down to ~16 days (Soszyński et al. 2017) might have light curves with alternating deep and shallow minima of light. For these targets, our algorithm for the periodicity search incorrectly assigned \( P_{\text{best}} \) to the time interval between two adjacent minima. However, this is half of the true, physical variability period. To detect the alternating minima and to better classify the variables, we visually inspected the light curves of all the PVS candidates at both \( P_{\text{best}} \) and \( 2 \cdot P_{\text{best}} \). We changed \( P_{\text{best}} \) into \( 2 \cdot P_{\text{best}} \) for those variables with clear signs of alternating minima.

4. Classification

The stars in our final sample have periods between ~0.38 and ~1240 days. In principle, they could be of any type among the pulsating stars, that is, RRLs, T2Cs, CCs, ACs, long-period variables (LPVs, which include Miras and SRVs, but not OSARGs), which have overly small amplitudes; Wray et al. 2004), non-pulsating variables (NPVs), which include, for example, EBs of any kind (detached, semi-detached, W UMa contact), and spotted stars, that is, stars which show variability due to the interplay of their rotation and the presence of large spots on their surface. Normally, these objects are associated to the pre-main sequence population. To discriminate between these types of variables, we adopted the criteria listed below. We point out that none of the following criteria were adopted alone for the classification, but they were all considered together. We list them in order of decreasing reliability.

Periods. Based on the literature, we adopted period thresholds for the quoted types of variables (see Table 1). We adopted conservative limits because the period thresholds of pulsating stars depend on many factors (metallicity, helium abundance, \( \alpha \) enhancement) and a detailed summary is beyond the aim of this paper. (1a) NPVs – These variables cover a wide range of periods, which includes all the periods of our targets. (1b) RRLs – Using OGLE data, Soszyński et al. (2014) found more than 38 000 RRLs, which is the largest homogeneous available catalog of RRLs. The period ranges of RRc and RRab are [0.20–0.54 days] and [0.28–1.00 days], respectively. These are quite usual thresholds, apart from the lower limit of RRab, for which ~0.4 is a more common value. However, since the shortest period in our list of PVSs is ~0.38 days, there is no difference in adopting one or the other value for the threshold. (1c) ACs – The most extensive catalog of bona fide ACs is provided by the OGLE survey Soszyński et al. (2015), for targets in the Large Magellanic Cloud (LMC) and Small Magellanic Cloud (SMC). These latter authors found ACs in a period range which is wider (0.38–2.7 days) than those quoted in theoretical studies (~0.5–2 days Fiorentino et al. 2006). We point out that the GC is an environment which in principle does not favor the formation of ACs, which arise either from very metal-poor or binary progenitors (Fiorentino & Monelli 2012). Nonetheless, to date, 20 of them were found in the Bulge (Soszyński et al. 2017). Therefore, we cannot discard a priori the possibility of finding ACs.

(1d) CCs – The sample of CCs with the widest range in period...
(0.25–208 days) is that of LMC and SMC CCS by OGLE (Soszyński et al. 2008a, 2010). We point out that the adopted empirical ranges are wider than theoretical predictions (Bono et al. 2000). (1e) T2Cs – The commonly accepted lower limit of the periods of T2Cs – which marks their separation from RRLs – is 1 day (Di Criscienzo et al. 2007). To set an upper threshold for T2Cs is a more delicate issue. RV Tau stars are the longest-period T2Cs and might show alternating deep and shallow minima. Therefore, two periods can be defined: the time passing between two deep minima (true period) and the time between two adjacent minima (formal period, by definition, half of the fundamental period). According to Wallerstein (2002), the maximum formal period is theoretically ~75 days. On the observational side, the longest formal periods in T2C catalogs are shorter than 90 days for Galactic globular clusters (Matsunaga et al. 2006; IRSF) and shorter than 80 days for the LMC, SMC, and Bulge (Soszyński et al. 2017, 2018; OGLE). However, it is not easy to detect the alternating minima, therefore it is not straightforward to discriminate between the two periods. Finally, we adopted a conservative threshold of 100 days for the formal period. (1) Miras – These variables are often classified based on their amplitude (A<sub>J</sub> > 0.8 mag) rather than their period. The most common lower period limit in the literature is 100 days (Ita & Matsunaga 2011; Catchpole et al. 2016). Nonetheless, Miras with shorter periods – down to ~80 days – were found in the Galactic Bulge (Soszyński et al. 2013) and furthermore Whitelock (2012) states that Miras could have periods shorter than 100 days. Therefore, we adopt 80 days as our threshold. We note that this is also the lower limit given in the General Catalog of Variable Stars (Samus’ et al. 2017) and the International Variable Stars Index definition of the class. On the other hand, the periods of Miras can be as long as 2000 days Whitelock (2012), which is longer than the longest period in our final sample of candidates.

**Bailey diagram.** We have collected the periods and Amp(K<sub>s</sub>) of RRLs, T2Cs, CCS, Miras and ACs in the literature, in the field of the Milky Way, the LMC, the SMC, the Galactic Globular Clusters and the Galactic Bulge (Matsunaga et al. 2006, 2013, 2016; Ripepi et al. 2014, 2015, 2016; Inno et al. 2015; Gran et al. 2016; Yuan et al. 2017). For each candidate PVS, we have checked its position in the Bailey diagram and compared this with the position of variables of known type, to narrow down the classification.

**Distance and position.** Based on known calibrations of PL relations for the different types of variables (see Table 1), for each target, we have calculated a set of provisional distances: d<sub>RRab</sub>, d<sub>Rc</sub>, d<sub>CCFO</sub>, d<sub>CCTF</sub>, d<sub>ACFO</sub>, d<sub>ACFU</sub>, d<sub>T2C</sub>, d<sub>Mira</sub>. These must be interpreted as the distance at which the target would be located if it was a variable of a given type, as in the subscript. Using the (l, b) coordinates, we also derive the provisional positions in Cartesian coordinates – x, y and z – of the targets in the Galaxy. We note that there are several caveats on distance. (3a) Metallicity dependence – At NIR wavelengths, the effect of metallicity on the PL relation of CCS is reduced compared to the optical (Marconi et al. 2010; Bhardwaj et al. 2015). The effect of metallicity is negligible also on the PLs of T2Cs (Di Criscienzo et al. 2007) and Miras (Whitelock et al. 2008). There is no solid theoretical or empirical evidence of metallicity dependence of the PL<sub>K</sub> of ACs (Fiorentino et al. 2006) either. On the other hand, the zero-point of the PL<sub>K</sub> relation of RRLs is affected by metallicity (Bono et al. 2001; Catelan et al. 2004; Marconi et al. 2015). We assume [Fe/H] = -1.0 (Sans Fuentes & De Ridder 2014; Hajdu et al. 2018) as a typical iron abundance of Bulge RRLs. We note that the error propagation associated to the uncertainty on [Fe/H] is negligible compared to the other factors. (3b) RVTs – There is currently debate over whether RVTs follow the PL of T2Cs (Matsunaga et al. 2006; Bhardwaj et al. 2017b) or not (Wallerstein 2002; Ripepi et al. 2015). More specifically, it is not clear how to separate intermediate-age RVTs from old, low-mass RVTs, which have a completely different evolution; it might be appropriate to give them a different name (e.g., V2342 Sgr stars, Catelan & Smith 2015). To sum up, the distances of even bona fide RVTs cannot be trusted. (3c) Reddening – We are aware that the pixels of the S14 reddening map are too large to

### Table 1. Criteria for the classification of variable stars.

| Type | P<sub>min</sub> (days) | P<sub>max</sub> (days) | Ref. | PLH (mag) | PLK<sub>f</sub> | Ref. |
|------|---------------------|---------------------|------|-----------|-------------|------|
| NPSs | <0.38               | >1000               | 1    | ...       | -1.56       | ...  |
| RRc(FO) | <0.38             | -0.54              | 2    | -0.971     | -0.987     | ...  |
| RRab(FU) | <0.38             | ~1                 | 2    | -2.226     | -2.250     | ...  |
| ACFO | <0.38              | ~1.2               | 3    | -2.42      | -3.54      | 12   |
| ACFU | ~0.50              | ~2.3               | 3    | -2.42      | -3.54      | 12   |
| CCF0 | <0.38              | ~4.5               | 4.5  | -2.851     | -2.890     | 13   |
| CCFU | ~0.80              | ~2.1               | 4.5  | -2.366     | -2.408     | 13   |
| T2C | ~1                 | ~100               | 6.7  | -3.585     | -3.418     | 14   |
| Miras | ~80                | ~2000              | 9.10 | ...       | -6.865     | 15   |

Notes. The zero points of the relations based on LMC variables were derived assuming, as the distance of the LMC, d<sub>LMC</sub> = 4.959 ± 0.09 ± 0.54 kpc (Pietrzyński et al. 2019). (a)/d) We did not derive distances of candidate T2Cs with P<sub>rest</sub> > 20 days (RVTs) because there is no general consensus on whether RVTs can be used as distance tracers or not (Matsunaga et al. 2006; Ripepi et al. 2015; Bhardwaj et al. 2017b; Braga et al. 2018a). (b) C-rich Miras with periods longer than ~5–350 days can be significantly affected by circumstellar reddening and do not obey the PL relation (Matsunaga et al. 2009; Ita & Matsunaga 2011). Since our color estimates are not accurate enough to separate C-rich and O-rich stars, we conclude that distance estimates for long-period Miras are not reliable.
provide accurate $E(J-K_s)$ estimates of targets within a region where the reddening pattern is so irregular at small scales. In Sect. 5.5, we discuss in more detail the comparison of reddening and distance with other works. We point out that for a fraction of targets our distance estimates might be overestimated due to underestimated extinction. However, NIR multi-band light curve templates are available for RRLs, CCs, and T2Cs (Braga et al. 2019; Inno et al. 2015 and Bhardwaj et al. 2017b), providing the possibility to estimate accurate $J$- and $H$-band mean magnitudes ($\langle J \rangle$, $\langle H \rangle$) and, in turn, to derive independent distance and $A_K$ estimates. This is the same approach used by M09 and M13, who did not use any reddening map. We point out that although we did derive $\langle J \rangle$, we only employed PL$_J$ and PL$_H$ relations to estimate extinction and distances, because differential reddening and variations of the reddening law have an overly large effect on targets in this region. The net effect of employing the PL$_J$ relations is to increase distances by ~0.4 kpcs, leading to a very unlikely peak of T2 distances around ~8.9 kpcs.

To estimate the $\langle J \rangle$ and $\langle H \rangle$ of RRLs and CCs, we used the light-curve templates of RRLs (Braga et al. 2019) and CCs (Inno et al. 2015). For T2Cs, only the $K_s$-band light-curve templates are available. However, they provide accurate mean magnitudes also when applied to the $H$-band data (Bhardwaj et al. 2017b), since the light-curve morphology in these two bands is very similar. Finally, we adopted the $J$-band CCFO light-curve template by (Inno et al. 2015) to derive $\langle J \rangle$ for T2Cs. We checked that, within the photometric errors, no difference is seen when using the phase correction (Bhardwaj et al. 2017a).

To perform the fits, one must reshape the amplitude using the $\text{Amp}(J)/\text{Amp}(K_s)$ and $\text{Amp}(H)/\text{Amp}(K_s)$ amplitude ratios. For RRLs and CCs, we adopted the ratios provided by Braga et al. (2018b) and Inno et al. (2015), respectively. For T2Cs, we checked that both $\text{Amp}(J)/\text{Amp}(K_s)$ and $\text{Amp}(H)/\text{Amp}(K_s)$ are equal to one, within the uncertainties, using globular cluster data from Matsunaga et al. (2006). Furthermore, we verified that minor phase shifts between the $J$, $H$, and $K_s$ light curves are negligible (~0.05 pulsation cycles). Therefore, the $J$- and $H$-band template fits were performed by leaving only the mean magnitude as a free parameter. Figure 5 shows the use of the template for an RRab (left), T2C (center), and CC (right). Computing $\langle J \rangle$ and $\langle H \rangle$ based on the intensity integral over the template provide mean magnitudes that are more accurate than the simple average of the two epochs, and this is even more critical when the $J$- or $H$-band data are sampled around the minimum.

We point out that due to the high extinction, not all of the candidate variables have $JH$-band data with acceptable photometric errors (<0.2 mag). This means that we could not apply the PL$_{JH}$ method to all of them. More precisely, we lack precise $J$-band data for 16 T2Cs and $H$-band data for 8 T2Cs and 1 CCFU.

To sum up, for 2 CCs, 5 RRabs, and 156 T2Cs, we derived two sets of distances and extinctions which we labeled $d_{s14}$, $A_{K_s(S14)}$, and $d_{pl}$, $A_{K_s(PL)}$, respectively, for those derived using the S14 reddening map and those obtained with the PL$_{JH}$ method.

**Velocity.** We use the provisional distances found before to derive provisional transverse velocities, using $v_t = 4.74\cdot d\cdot\mu$, where $d$ is in kiloparsecs, $\mu$ in mas/yr and $v_t$ in km/s. We note that to derive the direction and intensity of the tangential velocity, we summed the proper motion of Sgr A* (Reid & Brunthaler 2004) and our own proper motions. According to Braga et al. (2018a) we assume that this operation provides a proxy of absolute proper motions. We did not correct for the peculiar motion of the Sun because we are interested in understanding whether our targets follow the rotation curve of the Galactic disk, both on our side and on the far side of it. We assume, as a threshold on the reliability of $\mu$, a combined statistical error (CSE = $\sqrt{\sigma_{\mu}^2 + \sigma_{\mu}^2}$) of 2 mas yr$^{-1}$, as did Contreras Ramos et al. (2017), who used the same data. We also use the direction of $v_t$ as a diagnostic to discriminate between types of Cepheids belonging to different Galactic populations (e.g., we assume that the motion of CCs and young variables must be disk-like). We did not include radial velocities ($v_t$) in our analysis for two reasons. First, $v_t$ from the APOGEE survey (Majewski et al. 2017) are available only for seven of our targets. Second, and most important, with the exception of EBs and NPVs, our targets are radially pulsating stars, with radial-velocity amplitudes of the order of tens of kilometres per second (Bono et al. 2000; Feast et al. 2008) due to the inflation and deflation of the layers where absorption lines originate. Therefore, a single $v_t$ measurement is not an accurate estimate of the radial velocity of the barycenter of the star.

**Dereddened $H-K_s$ color.** Generally, our dereddened colors have large uncertainties (the peak of the distribution of $(H-K_s)_0$ is at ~0.3 mag), which also increase after the dereddening. Moreover, while we have accurate estimates of $(K_s)_0$, this is not true for other bands. Finally, since the reddening is extremely severe, we can only rely on the $H-K_s$ color, which covers a very small wavelength range. Nonetheless, $(H-K_s)_0$ comes in handy to

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**Fig. 5.** Left: phased $J$, $H$, and $K_s$-band light curves of a RRab over which we applied the light-curve template. The ID, type, and period of the star are labeled in red. Center: As in left, but for a T2C. Right: as in left, but for a CCFU.
Table 2. Properties of the candidate PVS.

| ID     | RA (deg) | Dec (deg) | Type | P (mag) | Y (mag) | T (mag) | H (mag) | V (mag) | S (mag) | T2C | ACFU | ACFO | RRab | RRc | SRV | Mira | NPV | DEB |
|--------|----------|-----------|------|---------|---------|---------|---------|---------|---------|-----|------|------|------|-----|-----|------|-----|-----|
| b333_59_98743 | 265.765256 | –28.941707 | 96 | 0.387960 | 20.132 | 0.114 | 18.622 | ±0.032 | 16.873 | ±0.032 |
| b333_46_54772 | 265.725409 | –29.200448 | 9 | 0.406797 | 20.797 | ±0.059 | 16.909 | ±0.015 | 14.403 | ±0.018 |
| b333_38_13973 | 266.087807 | –29.593169 | 469 | 0.505688 | . . . | 20.514 | ±0.159 | 18.547 | ±0.128 |
| b333_511_117170 | 266.345607 | –29.183243 | 95 | 0.515417 | . . . | 20.243 | ±0.294 | 17.946 | ±0.128 |
| b333_53_12695 | 265.773550 | –29.685386 | A5 | 0.524798 | . . . | . . . | 18.228 | ±0.047 | 17.550 | ±0.016 |

Notes. An asterisk indicates that the variable is most likely heavily saturated (Kₘ < 11 mag), therefore its magnitudes and amplitude might be inaccurate. A tentative classification (e.g., 96 in the fifth column in Table 2) from the average of the two available phase points or applying the template. The photometric and pulsation properties of Miras – except those in common with M09 – are not within this table and will be published in Nikzat et al. (in prep.). Only the first ten rows of the table are displayed. The full table is provided at the CDS.

Table 3. Kₕ-band time series of our candidate variables.

| Name      | HJD–2 400 000 | Mag | Err |
|-----------|---------------|-----|-----|
| b333_59_98743 | 55844.02155 | 14.704 | 0.043 |
| b333_59_98743 | 55423.14904 | 14.607 | 0.040 |
| b333_59_98743 | 55777.13660 | 14.865 | 0.054 |
| b333_59_98743 | 55778.19837 | 14.627 | 0.032 |
| b333_59_98743 | 55794.12621 | 14.674 | 0.047 |
| b333_59_98743 | 55806.15221 | 14.674 | 0.046 |
| b333_59_98743 | 55820.10395 | 14.705 | 0.045 |
| b333_59_98743 | 55830.02354 | 14.807 | 0.044 |
| b333_59_98743 | 55849.02315 | 14.518 | 0.036 |
| b333_59_98743 | 55987.37349 | 14.730 | 0.052 |

Notes. Column 1 gives the name, Col. 2 the heliocentric Julian day of the observation, Col. 3 the measured magnitude and Col. 4 the photometric error. We note that the light curves of Miras – except those in common with M09 – are not within this table and will be published in Nikzat et al. (in prep.). Only the first ten entries are listed. The full table is provided at the CDS.

Shape of the light curve. Although in the NIR many features of the light curves present in the optical are smoothed, it is still possible, by visually inspecting the light curves, to refine the classification of, for example, detached EBs (DEBs), which have narrow, alternating deep and shallow minima between almost flat plateaus.

We provide the light curves of these 1019 candidate variables in Table 3.

After taking account of the quoted criteria, we assigned a variable type to each of the candidate PVSs in the final sample. We ended up with unambiguous classification for 472 variables (5 RRab, 164 T2Cs, 3 CCFU, 1 ACFU, 16 SRVs, 210 Miras and 73 NPVs, of which 47 DEBs, see Table 2, second column). For the other 547 variables, we could not provide a solid classification due to uncertainties on the reddening and distance, and due to noisy light curves. More specifically, for 99 of them, the light curve is clear enough to be classified as “high-rank” candidate variables. However, due to the lack of other information, or ambiguity in the criteria listed above, we could only provide a tentative classification (e.g., 96 in the fifth column in Table 2 means that we are uncertain on whether b333_509_98743 is a RRc or a NPV). For the remaining 448 candidate PVSs, the light curve was not clear enough (due to noise, high photometric error, or secondary modulations), and therefore they were classified as “low-rank variables”. These too have to be considered as variables without a certain classification.

We point out that we have found the first bona fide AC towards the Galactic Bulge at such low latitudes. Its position in the Bailey diagram and the shape of its light curve were crucial to providing a solid classification. We have submitted a follow-up proposal to collect spectroscopic data with FIRE at Magellan, to estimate radial velocity and iron abundance for this star. Anomalous Cepheids are relatively rare objects, especially in the Milky Way, and spectroscopic information is available only for one of them (V19 in the globular cluster NGC 5466, McCarthy & Nemec 1997).

Figure 6 displays a sample of light curves of the bona fide T2Cs that we have detected. Figure 7 is the same for Miras and SRVs in common with M09.
Fig. 6. Phased $K_s$-band light curves of 60 variable stars classified as T2Cs. Periods are labeled at the bottom of each panel. The “b333” part of the name is omitted.
5. Matches with other catalogs

Although this work presents the most extensive survey of this region of the Galaxy for variables with periods longer than one day, and despite most of our targets being new detections, some were already found in similar investigations. In the following paragraphs, we compare our list with those found previously in the literature, more specifically, within the VSX catalog, the OGLE survey and – in order of increasing variable period (from RRLs, T2Cs, CCs to Miras) – CR18, M13, and M09. We did not find any match with the variables found by Dong et al. (2017) within the nuclear star cluster.

5.1. VSX and OGLE catalogs

We crossmatched our list of candidate variables with those in the VSX catalog and those found by the OGLE survey to see whether or not some of our variable candidates had already been found and/or classified. Indeed, we found nine matches within the eclipsing binaries catalog of the OGLE IV survey (Soszyński et al. 2011), the 16 T2Cs by M13, 7 eclipsing binaries catalog of the OGLE IV survey and – in order of increasing variable period (from RRLs, T2Cs, CCs to Miras) – CR18, M13, and M09. We did not find any match with the variables found by Dong et al. (2017) within the nuclear star cluster.

5.2. RR Lyr catalog of CR18

As discussed in Sect. 3.3.1, our periodicity search algorithm focused on periods longer than 1 day, and the variables with shorter periods are an incomplete sample. Nonetheless, our final catalog contains some candidate RRLs that are either uncertain or bona fide. We matched our catalog with the list of RRLs published by CR18 and found three sources in common. We note that we performed the match by ID because we used the same photometric data set. This means that there is no uncertainty associated to a cone search by coordinates.

The common sources are b333_616_55278, b333_614_37068 and b333_414_55144, which we classified as either OH stars or Miras; our classification for these stars is either “Mira?” or “Mira?”; thus matching the classification provided by the quoted studies. Concerning the matches with variables in the M09, Matsunaga et al. (2011), and M13 series of studies, more details are provided in Sects. 5.3 and 5.4.

5.3. T2C and CC catalog of M13

A previous search for Cepheids towards the GC was performed by M13 with the NIR SIRIUS camera at the IRSF 1.4m telescope. They found 45 variable stars, of which 20 were classified as Cepheids (16 T2Cs, 3 CCs and one generic Cepheid candidate).

The b333 tile overlaps completely with the IRSF survey sky area and we retrieved all the 20 Cepheids by M13 within our initial list of ~5 million sources. However, our periodicity search algorithm detected only 14 of them. Six Cepheids were not retrieved because their light curves have many uncertain phase points due to either saturation and/or blending. We manually extracted the light curves of the six missing Cepheids and, by knowing a priori their periods, we manually rejected the poor-quality data and obtained clean light curves for five of them. Unfortunately, b333_114_65025 is affected by severe blending and it was not possible to detect any periodic behavior even knowing the period a priori. We note that M13 also pointed out that this star (their #2) is in a very crowded field and that their photometry was not accurate.
The comparison with the M13 sample also provides a validation of our classification criteria. In fact, among the 14 variables that were retrieved automatically, we classified — without knowing, a priori, the classification by M13 — 13 of them as T2Cs and one as “T2C?”. We then checked the two samples and found that our classifications match those of M13, including their “Cept?”), which is our “T2C”.

We provide a more detailed discussion of the offsets in distance and reddening between our estimates and those of M13 in Sect. 5.5.

5.4. Miras catalogs of M09

Using the NIR SIRIUS camera at the IRSF 1.4 m telescope, M09 found 175 Miras towards the GC, and estimated their periods, distance, and extinction. They performed a sample selection based on the period, and marked only those with periods between 100 and 350 days as bona fide Miras, since they are the least affected by circumstellar reddening. We found a matching source for all their Miras in our initial photometric catalog. However, only 16 of them were retrieved in our final catalog of variable candidates. The reason for the low fraction of Miras retrieved in our final catalog is that M13 Miras are mostly located in the Bulge, therefore they are saturated — more severely than T2Cs — or are in the nonlinear regime of the VISTA images. A more complete discussion of these and other Mira variables in the VVV data will be presented in Nikzat et al. (in prep.).

Of these 16 variable candidates, 9 were classified as Miras, 6 as “Mira?”, and one as an SRV. We point out that the offsets of these and other Mira variables in the VVV data will be presented in Nikzat et al. (in prep.).

We provide a more detailed discussion of the offsets in distance and reddening between our estimates and those of M09 in Sect. 5.5.

5.5. A caveat on extinction and distance

In principle, the offsets in distance are not to be fully ascribed to offsets in reddening. As mentioned in Sect. 3.1, IRSF and VISTA have different diameters (1.4 m vs. 4 m). This means that stars in the linear photon count regime of the SIRIUS camera on IRSF can be saturated in VIRCAM. Figure 8 displays the difference in mean magnitude for the common sources.

Within the dispersion, which is very large (~0.1 mag), the average is zero. However, somewhat surprisingly, both the offset and the dispersion are larger for nonsaturated VVV targets than for saturated targets. The large individual offsets might be due to the different pixel scales of the two cameras: 0.25 mas pix$^{-1}$ for VIRCAM and 0.45 mas pix$^{-1}$ for SIRIUS. In such a crowded region, VIRCAM has the advantage of a spatial resolution that is almost twice as great as that of SIRIUS, meaning that it is more capable of resolving neighboring stars. Overall, we can conclude that the different saturation levels and pixel scales do not significantly affect the mean magnitudes, but the effect on individual stars is large, up to ±0.2 mag, meaning a relative systematic offset of up to ~8% on the distance.

Once it had been found that the offsets in $K_s$ have a minor impact on distances, we inspected the effects of the offsets in extinction. As mentioned in Sect. 4, for targets like CCs and T2Cs, we derived two sets of distance and extinction. We compare the $A_{K_s,PLK}$ and $A_{K_s}$ with those from M09 and M13. The main difference is that they used the Nishiyama et al. (2006, 2009) reddening law, which provides a different $A_{K_s}/E(J-K_s)$ ratio from ours (0.499 versus the 0.428 average value in Alonso-Garcia et al. 2017). This means that for a homogeneous comparison we need to compare $E(J-K_s)$ and not $A_{K_s}$.

For a rigorous interpretation of Fig. 9 one should take into account that in Sect. 4 we have adopted the same PL $K_s$ used by M09 for Miras, but different PL relations for T2Cs and CCs. Therefore, Fig. 9 also shows our $E(J-K_s)$ estimates both using the PLs listed in Table 1 (circles in Fig. 9) and using the same relations as M09 and M13, converted to the VISTA photometric system (crosses in Fig. 9). We note that to adopt the different calibrations of the PLs does not qualitatively change the observed trends. Furthermore, using the S14 map (left panel), our $E(J-K_s)$ estimates are always smaller than 5 mag, while those by M09 and M13 can be as large as ~7.5 mag and ~5.5 mag, respectively. These offsets generate the displayed trend: using the PLHK$_s$ solution for distance and $A_{K_s}$, the trend does not show up and the average offset is zero, within the standard deviation. Unfortunately, not all the common targets have a photometric PSF solution in our $H$ band images, since they are extremely faint, and therefore the sample is smaller. The fraction of targets with a PSF solution in the $J$ band drops dramatically, and it is pointless to perform the same analysis on them. We remind the reader that we did not adopt this technique for Miras (see Sect. 4, point 3c).

Most probably, the resolution of the S14 reddening map (6 × 6 arcmin) is insufficient to reproduce the fine details of the reddening patterns towards the GC, and the average values seem to be, overall, underestimated. One might expect that adopting the G12 reddening map, which has a resolution of 2 arcmin, would diminish this effect, but the average offset is negligible and the individual distances are even less reliable. In fact, using the G12 map, T2Cs display a very unlikely double-peaked distance distribution.

These offsets in reddening translate into significant distance offsets, which become dramatic for Miras, as displayed in Fig. 10.
Miras. The difference in reddening – and, in turn, on distances – is, on average, larger for Miras than for T2Cs and CCs. M09 estimated the distances to these targets to be smaller than 10 kpc for all of them. Most of them are located at between 7 and 10 kpc, and their peak is around 8.5 kpc, indicating that they belong to the Bulge. On the other hand, our distances are all larger than 10 kpc and their spatial distribution is almost constant over the 10–25 kpc distance range. One might invoke circumstellar extinction – which is accounted for in M09 extinction estimates, but not in the S14 map – to justify these offsets. However, as already mentioned, Miras with such short periods (<350 days) should not be significantly affected by circumstellar extinction (Ita & Matsunaga 2011). This means that, while possibly present, circumstellar extinction cannot, by itself, account for the whole offset. As mentioned above, we did not adopt the PLHKs solution for Miras.

CCs. These three Cepheids were first found by Matsunaga et al. (2011), who claimed that they are Bulge stars, and the first CCs ever found close to the GC. However, our different extinctions provide larger distances, which put them on the other side of the thin disk, at more than 9 kpc, both for $d_{S14}$ and $d_{PL}$, with small differences ($\lesssim 0.4$ kpc) between the two distance sets. We point out that, although these targets are saturated in the VVV, the difference between our $\langle K_s \rangle$ and those by M13 are small (–0.056 mag, 0.021 mag and –0.063 mag, for b333_215_87770, b333_215_86376 and b333_515_65036, respectively, where negative values indicate brighter $\langle K_s \rangle$ in our catalog). This means that the distance offsets displayed in the middle panel of Fig. 10 are not caused by a poor estimate of their mean magnitude, but should be ascribed to the difference in extinction only. We note that the adoption of the same PL as M13 affects the distances at a level of $\lesssim 0.3$ kpc for the $d_{S14}$ sample – which is negligible compared to our errors – and $\sim 1$ kpc for $d_{PL}$.

T2Cs. For all but three targets, we obtain a $d_{S14}$ that is larger by 1–2 kpc with respect to M13 distances, although a few objects
show distances that are similar or smaller (b333_215_69147, which is M13 #29). On the other hand, which is M13 #19; b333_416_9672, which is M13 #34; show distances that are similar or smaller (b333_215_69147, which is M13 #29). On the other hand, distances by a Gaussian fit to the rescaled distribution. The abscissa of the peak and the rescaled distance distribution (see text). The red solid line shows a Gaussian fit to the rescaled distribution. The absissa of the peak and the rescaled distance distribution (see text).

Fig. 11. Top: light blue dashed histogram showing the $d_{S14}$ distribution of BLHs and WVs within 20 kpc. The dark blue histogram shows the rescaled distance distribution (see text). The red solid line shows a Gaussian fit to the rescaled distribution. The absissa of the peak and the sigma of the Gaussian are labeled. Bottom: as in top, but for $d_{PL}$.

To sum up, the effect on the distance estimates of the offsets in $A_K$ is on average that of shifting targets to higher distances. However, at least for T2Cs, using the PLHK$_s$ method provides results that are similar to those of M13. Since most of these targets are saturated, their proper motions have large uncertainties, and therefore we could not decipher which set of $A_K$-distances is correct based on their proper motions (e.g., the three CCs should have disk kinematics if our distance estimates are correct).

Therefore, we inspect the $d_{S14}$ and $d_{PL}$ distribution of T2Cs in greater detail; they are displayed as light blue histograms in Fig. 13. The plain distributions provide biased information. In fact, one must take into account the depth effect on the distance estimates of the objects. This is consistent with the spatial distribution of RRLs found by CR18, that is, they are located around the GC but their sky surface density decreases towards the GC itself.

Fig. 12. The T2Cs (blue diamonds) are uniformly distributed enclosing distances between 7.5 and 8.0 kpc is undersampled in $d_{S14}$, making the distribution fairly asymmetrical. This is further evidence that the S14 map does not provide extinctions that are systematically offset by a fixed amount, but rather shows that there are individual offsets in $A_K$. On the other hand, the distances obtained using the PLHK$_s$ method are more reliable, because the peak of Gaussian is closer to the estimates in the literature. Finally, the smaller HWHM of the $d_{PL}$ distribution goes in the correct direction. In fact, although investigations of the spatial distribution of RRLs and T2Cs towards the entire Bulge generally provide larger HWHM ($\sim$0.8–1.0 kpc Pietrukowicz et al. 2015; Bhardwaj et al. 2017a), our investigation is more limited in sky area, and therefore we expect a tighter distribution of T2Cs.

Finally, we discuss the huge effect played by the reddening law. Not only, as already mentioned, is the total-to-selective absorption ratio different between Nishiyama et al. (2006, 2009) and Alonso-García et al. (2017), but a macroscopic effect is played by the extinction ratio $A_J/A_K$, which is 1.731 and 1.88 for the two reddening laws, respectively. Of course the photometric systems are different, and therefore the comparison is not straightforward. Nonetheless, the PLHK$_s$ method strongly depends on this ratio and the effect is a systematic decrease of the distances by $\sim$1.2 kpc when using the Nishiyama et al. (2006, 2009) reddening law.

All of these considerations provide strong evidence that further insight into the reddening law towards the central-most regions of the Galaxy is needed in order to correctly interpret the results of past and future observations, especially due to the uncertainty on distances and differential reddening effects.

6. Catalog properties

The positions of the 1019 PVSs in the Galaxy are displayed in Fig. 12. The T2Cs (blue diamonds) are uniformly distributed within the b333 tile. This is expected, since T2Cs are mostly old, with a fraction of intermediate-age stars (Wallerstein 2002; Iwanek et al. 2018).

We also point out that the position of RRLs, which are found only close to the borders of the tile, is due to the intrinsic faintness of these objects. This is consistent with the spatial distribution of RRLs found by CR18, that is, they are located around the GC but their sky surface density decreases towards the GC itself (see their Fig. 6).

The distribution of the periods of our variables, displayed in Fig. 13, outlines some interesting features. First of all, we note a lack of variables in the interval between $\sim$20 and $\sim$365 days. The reason for this lack of stars is due to the aliases at one year and at half a year, and to the transition between WVs and RVTs: the light curves of RVTs might display alternating deep and shallow minima, and are therefore more difficult to classify.

Six of the nineteen RVTs display clear alternating minima. We have inspected their position and conclude that they are not peculiar with respect to the other RVTs and T2Cs in general. (Soszyński et al. 2017) found that WVs with periods longer than 16 days also might show alternating minima, but we did not find any of these.

Figure 13 also displays a lack of variables with solid classification in the period between 1 and 2 days, despite the large number of variables (119) detected at these periods. This is mainly due to the overlap of many variable types (T2Cs, CCFOs, CCFUs, ACFOs, ACFUs) in that region of the Bailey diagram, which is one of our main diagnostics for the classification.
Fig. 12. Map in Galactic longitude and latitude ($l$, $b$) showing the distribution in the sky of the sample variable stars. Map obtained with Aladin Sky Atlas v10.076 (Bonnarel et al. 2000), using the 2MASS $JHK_s$ images catalog (Skrutskie et al. 2006). Blue diamonds represent T2Cs, which are spread over the whole field. Gray diamonds display Miras, which are more concentrated at low latitudes. Red ovals display CCs (the ones by Matsunaga et al. 2011). The magenta squares in this map are the RRLs, the cyan square is the AC, and the small white plus symbols represent all the other types of variables.

Fig. 13. Period distribution of our final list of PVSs. The dashed lines mark the aliases at 1 day and 1 year. The black histogram includes all the 1019 candidate PVSs found, the dashed magenta, blue, and red histograms represent RRab, T2C, and CCFU, respectively. The filled light blue bin at $P \lesssim 2$ days represents the ACFU.

Fig. 14. Bailey diagram of the PVS in our final list. Small black plus symbols mark the position of stars with uncertain classification. Blue diamonds mark T2Cs, magenta triangles mark RRab stars, red circles mark CCFUs, and the cyan square marks the ACFU. The solid lines represent the Oosterhoff I (left) and Oosterhoff II (right) sequences of RRab stars (Cacciari et al. 2005), rescaled using $\text{Amp}(K_s)/AB$ amplitude ratios (Braga et al. 2018b).
According to our estimates, its distance is 14 ± 0.7 kpc. Also taking into account systematic uncertainties due to reddening, there is no way that this ACFU can be within the Bulge. We derived a height above the Galactic plane of only 83 ± 4 pc. Moreover, its proper motion has an almost null latitudinal component (Fig. 16, middle panel). However, its tangential velocity is too small (109.2 ± 99.4 km s\(^{-1}\)) to be consistent with thin disk dynamics, and therefore it must be either a halo or a thick disk star. Spectroscopic follow-ups of the target are planned to constrain its chemical and dynamical properties. We have already submitted a proposal for observations with FIRE at Magellan. These will be useful to better constrain the complicated evolutionary picture of ACs, which is, to date, uncertain (Fiorentino & Monelli 2012; Iwanek et al. 2018).

7. Conclusions

We present the most extensive catalog of variable stars in the region of the GC, which includes 1019 objects. The catalog contains accurate coordinates, NIR (ZyJH) magnitudes, \(K_s\)-band mean magnitudes, \(K_s\)-band light amplitudes, and periods. We also provide proper motions for 530 targets and extinctions \(A_K\), plus individual distances for 220 targets. For 472 variables, we provide a high-rank, unambiguous classification. The latter sample includes 5 RRab, 164 T2Cs, 3 CCFU, 1 ACFU, 16 SRVs, 210 Miras, and 73 NPs (of which 47 DEBs).

For all these candidate variables, we show the NIR CMDs, proper motions, period distributions, and spatial distributions for different types of variables. Including the most important distance indicators (RRLs, Cepheids).

T2Cs. 164 bona fide T2Cs, among which 45 BLHs, 100 WVs, and 19 RVTs were retrieved. Type II Cepheids are uniformly distributed across the b333 field of view. It was not trivial to provide a solid classification for variables with periods between 1 and 2 days because in that period range many possible types of variability overlap (T2Cs, CCFUs, CCFOs, ACFUs, and all types of eclipsing binaries). We have looked in detail into their individual reddening estimates – both individual and using extinction maps – and we investigated their distance distribution. We also compared our results with those by Matsunaga et al. (2013) since there are 16 T2Cs in common, but all we can conclude is that our distances are 1–2 kpc larger than those found in the literature. This is a very large offset, but this should not be surprising. In fact, we estimate that a ~10% offset in the \(A_H/A_K\) extinction ratio causes a shift as large as ~1.2 kpc in the estimate for a star at the distance of the GC, using the same calibrating relations, \(E(J − K_s)\), mean magnitudes, and photometric system. This is a stark warning for all works focusing on distance estimates in regions that are close to the GC.

Miras. We retrieved 210 Miras candidates with periods between 87 and 943 days. We did not publish their astrophysical properties (coordinates, periods, mean magnitudes, etc.) because this is the aim of Nikzat et al. (in prep.). Nonetheless, we compared our photometric solution with that of Matsunaga et al. (2009). The effect of the uncertainties on reddening are dramatic for short-period (\(P < 350\) days) Miras and turn into differences in distance as large as 10 kpc. Invoking circumstellar extinction is not enough to justify these differences, and both the amount of reddening and the effect in the adopted reddening law (Nishiyama et al. 2009; Alonso-Garcia et al. 2017) play a major role in contributing to systematic uncertainties.

CCs. Among the point-source catalog of our PSF photometry, we retrieved the three CCs found by Matsunaga et al. (2011).
Our periodicity search algorithm did not automatically detect them, since their light curves are relatively noisy. By adopting an a priori period, we were able to remove outliers from the light curves and fit their light curves to derive mean magnitudes and amplitudes.

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