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Meta-Analysis of the INSIG2 Association with Obesity Including 74,345 Individuals: Does Heterogeneity of Estimates Relate to Study Design?

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Abstract

The INSIG2 rs7566605 polymorphism was identified for obesity (BMI$\geq 30$ kg/m$^2$) in one of the first genome-wide association studies, but replications were inconsistent. We collected statistics from 34 studies (n = 74,345), including general population (GP) studies, population-based studies with subjects selected for conditions related to a better health status ('healthy population', HP), and obesity studies (OB). We tested five hypotheses to explore potential sources of heterogeneity. The meta-analysis of 27 studies on Caucasian adults (n = 66,213) combining the different study designs did not support overall association of the CC-genotype with obesity, yielding an odds ratio (OR) of 1.05 (p-value = 0.27). The I$^2$ measure of 41% (p-value = 0.015) indicated between-study heterogeneity. Restricting to GP studies resulted in a declined I$^2$ measure of 11% (p-value = 0.33) and an OR of 1.10 (p-value = 0.015). Regarding the five hypotheses, we showed that (a) some difference between GP and HP studies (p-value = 0.012) and (b) an association in extreme comparisons (BMI$\geq 32.5, 35.0, 37.5, 40.0$ kg/m$^2$ versus BMI$< 25$ kg/m$^2$) yielding ORs of 1.16, 1.18, 1.22, or 1.27 (p-values 0.001 to 0.003), which was also underscored by significantly increased CC-genotype frequencies across BMI categories (10.4% to 12.5%; p-value for trend = 0.0002). We did not find evidence for differential ORs (c) among studies with higher than average obesity prevalence compared to lower, (d) among studies with BMI assessment after the year 2000 compared to those before, or (e) among studies from older populations compared to younger. Analysis of non-Caucasian adults (n = 4889) or children (n = 3243) yielded ORs of 1.01 (p-value = 0.94) or 1.15 (p-value = 0.22), respectively. There was no evidence for overall association of the rs7566605 polymorphism with obesity. Our data suggested an association with extreme degrees of obesity, and consequently heterogeneous effects from different study designs may mask an underlying association when unaccounted for. The importance of study design might be under-recognized in gene discovery and association replication so far.

Introduction

One of the first genome-wide association (GWA) studies ever and the first on obesity identified the INSIG2 gene represented by the rs7566605 polymorphism as a novel gene for common obesity [1]. Functional evidence depicted the INSIG2 gene from the very start as an interesting candidate for obesity as being involved in the reversed cholesterol transport by an interaction with sterol regulatory element-binding proteins (SREBPs) [2], which are transcription factors that activate the synthesis of cholesterol and fatty acids in the liver and other organs [3].

The observed SNP-obesity-association was replicated in some, but not in all studies [4–11]. A letter to Science by the authors of the initial report in response to the emerging debate of the early inconsistent results [1] raised the question of whether the association might be more pronounced in studies that were not ascertained for reasons related to better health status, when comparing more severely obese subjects with normal controls, in populations with a higher prevalence of obesity or in populations with a higher mean age. The need for a meta-analysis was stated there for the first time and re-stated by Lyon and colleagues [12]. Furthermore, a secular trend for increasing prevalence of obesity was observed in two large population-based studies from the same geographical region using the same protocols but one recruited 1994/95 (KORA-S3) and the other 1999–2001 (KORA-S4) [13]. The later study showed a stronger INSIG2-obesity-association compared to the earlier study: This raised the additional question whether the changes in nutritional intake and physical activity [14,15] believed to contribute to the increase in the prevalence of obesity during the last decades were the reason for some of the between-study heterogeneity observed for this SNP’s association with obesity.

The inconsistent reported associations and the many resulting debates motivated us to undertake a systematic meta-analysis of all available data to investigate potential causes of heterogeneity and to look for consistent results among subgroups. It was thus the specific aim of this meta-analysis to explore five hypotheses for heterogeneity of the rs7566605 association with obesity: (Hypothesis 1) The association depends on study design. Therefore, we classified studies as general population-based (GP) when they were neither selected for any disease nor for not having any disease, as any selection of this type was shown to potentially induce bias for outcomes associated with the disease [16]. We classified studies as ‘healthy population’ (HP) when they were selected for reasons related to a better health status (i.e. studies including subjects from working populations or studies excluding subjects for with diseases such as diagnosed type 2 diabetes), or obesity studies (OB) when they were specifically designed to investigate obesity, usually case-
control or family studies. We did not include studies that were ascertained for any disease to reduce overly complexity. (Hypothesis 2) The association is more pronounced when comparing more extreme cases of obesity with normal or lean subjects, or (Hypothesis 3) among studies with a greater percentage of obese individuals. (Hypothesis 4) The association is differentially seen in studies including subjects with a higher age compared to studies based on younger populations, or (Hypothesis 5) more pronounced in studies with a more recent assessment of BMI (after year 2000) assuming that these studies would reflect the changes in dietary habits and physical activity of the last decade and assuming that subjects with the INSIG2 risk genotype are more prone to gain weight in such an environment.

### Results

#### Data Collected

We have gathered data on 34 studies from across Europe, North America, and Asia that met the inclusion criteria including a total of 74,345 subjects. All studies were categorized a priori according to their study design (GP, HP, OB), ethnicity, and whether it was an adult or children population (Tables S1 and S2). A more detailed description can be found in the original study publications [1,4–12,17–20].

#### Main Analysis

The main results are summarized in Table 1: The meta-analysis of 27 studies of Caucasian adults (n = 66,213) showed a fixed effect odds ratio (OR) of 1.076 (p-value = 0.023) and random effects OR of 1.051 (p-value = 0.028) with the I² measure indicating significant between-study heterogeneity (I² = 41.0%, 95% confidence interval, CI = [6.6%, 62.8%], Q-test p-value = 0.015).

#### Some Heterogeneity Might Relate to Study Type

The I² measure declined (I² = 10.9%, 95% CI = [0.0%, 48.1%], Q-test p-value = 0.329) by restricting to the 16 GP studies, including 48,844 subjects, yielding a fixed [random] effects OR of 1.097 [1.092] (p-value = 0.015 [0.035]). The OR estimates were similar when excluding early published studies (p = 0.054 [0.060]).

The I² measure for heterogeneity was zero among the five HP studies, including 7640 subjects, and combined estimates yielded a tendency towards a protective OR of 0.796 (p-value = 0.028) without remarkable change when excluding the early published studies.

The OR among the six OB studies, including 9729 subjects, of 1.152 using the random effect model was higher than the one for the GP studies, but not statistically significantly different from unity (p-value = 0.253). There was substantial heterogeneity among the OB studies (I² = 63.2%, Q-test p-value = 0.018). The OR estimates were similar when excluding the early published studies.

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**Table 1. Main meta-analysis results of the INSIG2 rs7566605 association with obesity.**

| # cases/controls (# studies) | OR (p-value) fixed effect | OR (p-value) random effect | I² (p-value) | P-value test for difference of fixed [random] effect ORsb |
|-------------------------------|---------------------------|---------------------------|-------------|--------------------------------------------------------|
| **All-CA**                    |                           |                           |             |                                                        |
| 16,365/49,848 (27)            | 1.076 (0.023)             | 1.051 (0.268)             | 41.0 (0.015)|                                                        |
| **All-CA**                    |                           |                           |             |                                                        |
| 10,761/30,168 (19)            | 1.077 (0.074)             | 1.054 (0.361)             | 36.4 (0.057)|                                                        |
| **GP**                        |                           |                           |             |                                                        |
| 9162/39,682 (16)              | 1.097 (0.015)             | 1.092 (0.035)             | 10.9 (0.329)| GP vs HP: 0.004 [0.005]                                  |
| **GP**                        |                           |                           |             |                                                        |
| 5803/23,243 (12)              | 1.100 (0.054)             | 1.080 (0.060)             | 1.3 (0.431) | GP* vs HP*: 0.038 [0.089]                               |
| **HP**                        |                           |                           |             |                                                        |
| 1307/6333 (5)                 | 0.796 (0.028)             | 0.796 (0.028)             | 0.0 (0.415) | HP vs OB: 0.038 [0.022]                                 |
| **HP**                        |                           |                           |             |                                                        |
| 795/4116 (3)                  | 0.821 (0.135)             | 0.772 (0.198)             | 46.2 (0.156) | HP* vs OB*: 0.038 [0.109]                               |
| **OB**                        |                           |                           |             |                                                        |
| 5896/3833 (6)                 | 1.163 (0.018)             | 1.152 (0.253)             | 63.2 (0.018) | OB vs GP: 0.478 [0.680]                                 |
| **OB**                        |                           |                           |             |                                                        |
| 3119/2338 (4)                 | 1.155 (0.144)             | 1.179 (0.337)             | 65.0 (0.036) | OB* vs GP*: 0.657 [0.693]                               |
| **ALL-NC**                    |                           |                           |             |                                                        |
| 553/4336 (4)                  | 1.013 (0.936)             | 0.945 (0.818)             | 43.5 (0.150) |                                                        |
| **ALL-CH**                    |                           |                           |             |                                                        |
| 1802/1441 (3)                 | 1.147 (0.216)             | 1.147 (0.216)             | 0.0 (0.830)  |                                                        |

Pooled association for all Caucasian adult studies combined (All-CA) as well as stratified by study type (GP = general population, HP = healthy population, OB = obesity study), for all non-Caucasian studies (All-NC), or the children studies (All-CH) indicating differential results among GP or HP studies (Hypothesis 1). Numbers stated are the number of obese (BMI≥30 kg/m², ‘cases’) and non-obese (BMI<30 kg/m², ‘controls’) subjects with the number of studies, the recessive model ORs (p-value) comparing the odds for obesity among the subjects with the CC genotype versus subjects with the CG or GG genotype using fixed or random effects models, the I² measure (p-value of the Q-statistics), and the p-value testing for pair-wise difference between GP, HP, and OB studies.

aExcluding studies published before the response letter in Science by Herbert et al., December 2006 [1], in which the hypothesis of potential heterogeneity due to study design and a first call for this meta-analysis were stated (i.e., excluding American, Polish, NHS, KORA_S4, Essex_trios, EPIC_Norfolk, MRC_Ely, DESIR, SHIP, OB_adult).

bP-values corrected for pair-wise comparison of three subgroups need to be multiplied by three.

OR = Odds Ratios.

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Figure 1A–1C shows forest plots of the Caucasian adult studies with combined estimates by study type. The p-value testing for difference between the GP and the HP combined estimates was 0.004 [0.012 when corrected for pair-wise testing of three subgroups] and 0.039 [0.089] when excluding the early published studies. Thus, there is some, but not completely conclusive evidence for Hypothesis 1 that study design might explain some of the between-study heterogeneity of this genetic association.

### Studies in Non-Caucasian Adults or Children

In the pooled analysis of the four studies on non-Caucasian adults (n = 4859), we found no significant association of the CC genotype with obesity. The pooled analysis of the three pediatric studies (n = 3243) was also not significant (Table 1, Figure 1D and 1E).

### Increasing Strength of Association in More Extreme Comparisons

Our data suggested an association with increased ORs among the Caucasian adults when more extremely obese subjects were compared to lean controls (Table S3; Hypothesis 2); combining over all studies, the ORs gradually increased from 1.156 to 1.183, 1.221, or 1.265 (p-values ranging from 0.001 to 0.003) when moving the BMI cut-off for the obese cases from 32.5 to 35.0, 37.5 or 40.0 kg/m², respectively, and comparing to controls with BMI <25 kg/m². A similar trend for increasing ORs was seen when comparing extremely obese subjects against controls with BMI <30 kg/m² or BMI <20.0 kg/m². Among the GP studies, the ORs increased from 1.198 to 1.257, 1.313, or 1.414 (p-values ranging from 0.001 to 0.003), respectively. The analogous comparison for HP studies revealed that the protective influence of the CC homozygous was reversed in the more extreme comparisons with ORs from 0.856 to 0.959, 1.139, or 1.604. For the OB studies, the ORs of the analogous comparisons were well above unity for all comparisons, but did not show a trend.

As summarized in Table 2, the accompanying trend analyses of the CG genotype frequencies across the various BMI categories indicated significantly increased CC genotype frequencies from 10.4% to 12.5% (p-value testing for trend = 0.0002), which persisted when excluding the early published studies (p-value = 0.0008). A similar trend was seen among GP studies.

In both types of analyses, the varying cut-point ORs as well as the trend in genotype frequencies by BMI categories, suggest an association of the rs7566605 with extreme obesity compared to normal controls (Hypothesis 2).

### Further Hypotheses on Potential Sources of Heterogeneity

Table 3 summarizes the results of the further three hypotheses to explain heterogeneity, which were tested for the Caucasian adult studies. Hypothesis 3: there was some tendency towards higher ORs among ‘more obese’ study populations compared to the ‘less obese’ study populations (p-value = 0.052 [0.285] testing for difference of the fixed [random] effects ORs), but not statistically significant. Hypothesis 4: there was no evidence for any difference between studies from older populations (i.e. mean age of subjects above 50 years) as compared to studies from younger populations (i.e. mean age below 50 years). Hypothesis 5: there was a tendency towards more pronounced ORs for the studies with BMI assessed after the year 2000 as compared to studies with BMI assessed before 2000 (p-value = 0.007 [0.095]), but not statistically significant given the various tests performed and particularly not when excluding the early published studies (p-value = 0.086 [0.248]). Hypotheses 3–5 were not tested in the HP or OB stratified analyses as too few (3–6) studies were available.

### Sensitivity Analyses

The sensitivity analyses (Table S4) indicated robustness of estimates towards selection of gender or age and no significant difference between published and unpublished studies. Excluding the two studies with self-reported BMI from the overall GP meta-analysis resulted in a slight increase of the OR estimate.

### Genetic Model

We specifically examined the association under the recessive genetic model as suggested by the original paper [1]. Our data on the raw numbers of obese or non-obese subjects with one of the three genotypes underscored a recessive model in the Caucasian adult studies combined (OR_CCversusGG = 1.112 [1.029, 1.203], p-value = 0.007, and OR_CGversusGG = 0.988 [0.940, 1.037], p-value = 0.618) as well as among the GP studies (OR_CCversusGG = 1.076 [0.976, 1.183], p-value = 0.142, and OR_GGversusGG = 0.995 [0.936, 1.036], p-value = 0.813).

### Secondary Analyses on BMI as a Quantitative Outcome

The secondary analyses on BMI as a quantitative outcome were only performed in GP and HP studies. These analyses generally showed results consistent with the obesity analyses, but less, if any, significance (Figure S1, Tables S3, S6, and S7).

### Discussion

We conducted a systematic meta-analysis on published and unpublished studies by collecting summary statistics on the association of the rs7566605 SNP near the INSIG2 gene using a recessive genetic model as proposed by Herbert and colleagues. This SNP was highly debated and the inconsistent findings were very much puzzling underscored by again inconclusive findings in two recent publications [21,22]. To solve this puzzle, we collected aggregated study-specific data from 34 studies with a total of 74,345 subjects analyzed according to a standardized model.

The main analysis did not support evidence for an overall association with obesity of the CC-genotype compared to the CG/GG (OR = 1.05, p-value = 0.268). Our data suggested an association for more extreme obese subjects (BMI ≥32.5, 35.0, 37.5, 40.0 kg/m²) compared to normal controls (BMI <25 kg/m²) with ORs increasing to 1.16, 1.18, 1.22, 1.27 (p-values between 0.001 and 0.003) and significantly increased CC-genotype frequencies with increasingly high BMI categories (10.4% to 12.5%, p-value for trend = 0.0002). The main analysis pointed towards significant between-study heterogeneity with an I² measure of 41%. When we restricted the analysis to GP studies, the I² declined to 11% and the OR increased to 1.10. This is in-line with a very recently published study, which found the OR to increase from 1.02 for a combined analysis of diverse types of studies including 16,781 subjects to an OR of 1.15 when restricting to the general population-based INTER99 cohort including 6,158 subjects [21]. This was the largest GP study on this SNP-association prior to this meta-analysis.

### The Degree of Obesity and the Study Design as a Potential Source of Heterogeneity

The results of our analyses suggest an association of this SNP with extremely obese subjects compared to normal controls, but future research will need to confirm this finding. Study design can impact how many extremely obese subjects are included in the

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**Table 3**

| Hypotheses | Description |
|------------|-------------|
| Hypothesis 1 | Study design might explain some of the between-study heterogeneity of this genetic association. |
| Hypothesis 2 | Comparisons |
| Hypothesis 3 | Increasing Strength of Association in More Extreme Comparisons |
| Hypothesis 4 | Further Hypotheses on Potential Sources of Heterogeneity |
| Hypothesis 5 | The Degree of Obesity and the Study Design as a Potential Source of Heterogeneity |

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**Table 2**

Comparing extremely obese subjects against controls with BMI <30 kg/m² or BMI <20.0 kg/m². Among the GP studies, the ORs increased from 1.198 to 1.257, 1.313, or 1.414 (p-values ranging from 0.001 to 0.003), respectively. The analogous comparison for HP studies revealed that the protective influence of the CC homozygous was reversed in the more extreme comparisons with ORs from 0.856 to 0.959, 1.139, or 1.604. For the OB studies, the ORs of the analogous comparisons were well above unity for all comparisons, but did not show a trend.
### Table: Odds Ratio estimates for INSIG2 rs7566605 association with obesity

| Study                        | OR [95% CI] | Weight [%] | OR [95% CI] |
|------------------------------|-------------|------------|-------------|
| **A. Obesity analysis for GP adult studies** |             |            |             |
| Clinto_genetic               | 0.42        | 2.12 [0.67, 6.68] |             |
| SHARE_Caucasian              | 0.44        | 1.11 [0.36, 3.42] |             |
| Kiel_ageing                  | 0.60        | 1.30 [0.49, 3.43] |             |
| Kiel_genetics                | 1.43        | 0.67 [0.36, 1.26] |             |
| QFS                          | 1.82        | 0.81 [0.46, 1.40] |             |
| FHS_UNREL                    | 2.48        | 1.28 [0.79, 2.03] |             |
| EPIC_Potsdam                 | 4.58        | 1.04 [0.73, 1.47] |             |
| NFBG_1966                    | 5.90        | 0.99 [0.72, 1.36] |             |
| Czech_MONICA                 | 7.04        | 1.20 [0.90, 1.59] |             |
| KORA_S3                      | 8.56        | 0.91 [0.70, 1.17] |             |
| DESIR                        | 8.91        | 0.91 [0.71, 1.17] |             |
| EPIC_Norfolk                 | 10.18       | 1.00 [0.79, 1.26] |             |
| SHIP                         | 10.26       | 1.15 [0.91, 1.45] |             |
| Collus                       | 10.56       | 1.09 [0.86, 1.37] |             |
| KORA_S4                      | 10.69       | 1.33 [1.06, 1.67] |             |
| DECODE                       | 16.41       | 1.29 [1.07, 1.56] |             |
| **Pooled FE OR [95% CI]**    | 100.00      | 1.10 [1.02, 1.18] |             |
| Test for heterogeneity: $\chi^2 = 16.83, df = 15 (P = 0.38), I^2 = 10.9\%$ |             |            |             |
| Test for overall effect: $Z = 2.42$ (P = 0.02) |             |            |             |

| **B. Obesity analysis for HP adult studies** |             |            |             |
| HERITAGE_White               | 3.70        | 0.30 [0.10, 0.86] |             |
| NHS                          | 12.03       | 0.81 [0.45, 1.46] |             |
| MRC_Ely                      | 25.90       | 0.73 [0.49, 1.10] |             |
| SAPPHIR                      | 27.63       | 0.85 [0.58, 1.27] |             |
| NPHSII                       | 30.74       | 0.89 [0.62, 1.29] |             |
| **Pooled FE OR [95% CI]**    | 100.00      | 0.80 [0.65, 0.98] |             |
| Test for heterogeneity: $\chi^2 = 3.93, df = 4 (P = 0.41), I^2 = 0\%$ |             |            |             |
| Test for overall effect: $Z = 2.19$ (P = 0.03) |             |            |             |

| **C. Obesity analysis for OB adult studies** |             |            |             |
| Essen_obese                  | 10.99       | 1.64 [0.94, 2.89] |             |
| OBENUTIC                     | 16.21       | 1.09 [0.74, 1.60] |             |
| Utah_obese                   | 17.19       | 0.78 [0.55, 1.21] |             |
| OB_adult                     | 17.22       | 0.85 [0.60, 1.21] |             |
| Berlin_obese                 | 17.85       | 1.51 [1.08, 2.11] |             |
| American_Polish              | 20.55       | 1.40 [1.08, 1.83] |             |
| **Pooled RE OR [95% CI]**    | 100.00      | 1.15 [0.90, 1.47] |             |
| Test for heterogeneity: $\chi^2 = 13.58, df = 5 (P = 0.02), I^2 = 63.2\%$ |             |            |             |
| Test for overall effect: $Z = 1.14$ (P = 0.25) |             |            |             |

| **D. Obesity analysis for Non-Caucasian adult studies** |             |            |             |
| SHARE_South_Asian            | 6.04        | 0.82 [0.23, 2.89] |             |
| GPS                          | 10.95       | 0.40 [0.16, 1.00] |             |
| Asian_Indian                 | 21.72       | 1.47 [0.76, 2.85] |             |
| CLHNS                        | 61.29       | 1.07 [0.72, 1.59] |             |
| **Pooled FE OR [95% CI]**    | 100.00      | 1.01 [0.74, 1.38] |             |
| Test for heterogeneity: $\chi^2 = 5.31, df = 3 (P = 0.15), I^2 = 43.5\%$ |             |            |             |
| Test for overall effect: $Z = 0.08$ (P = 0.94) |             |            |             |

| **E. Obesity analysis for pediatric studies** |             |            |             |
| PIONEER                      | 5.02        | 1.43 [0.55, 3.77] |             |
| CHOBES                       | 27.57       | 1.22 [0.80, 1.84] |             |
| Essen_tricos                 | 67.42       | 1.10 [0.85, 1.43] |             |
| **Pooled FE OR [95% CI]**    | 100.00      | 1.15 [0.92, 1.42] |             |
| Test for heterogeneity: $\chi^2 = 0.37, df = 2 (P = 0.83), I^2 = 0\%$ |             |            |             |
| Test for overall effect: $Z = 1.24$ (P = 0.22) |             |            |             |

**Figure 1. Forest plot on Odds Ratio estimates of the INSIG2 rs7566605 association with obesity.** Association with obesity (BMI ≥ 30 kg/m²) for (A) the Caucasian adult general population-based (GP) studies, (B) the Caucasian adult studies ascertained for reasons related to health status (HP), (C) the Caucasian adult obesity case-control studies (OB), (D) the non-Caucasian adult studies, and (E) the pediatric studies. Shown are recessive model OR estimates comparing the CC genotype versus CG or GG for each study and pooled estimates. The fixed effect (FE) model OR is shown in case of no significant heterogeneity as tested by the Q-statistics; a random effect (RE) model is shown otherwise. doi:10.1371/journal.pgen.1000694.g001
study. Study designs that sample more extremely obese subjects will have greater power to detect the association, while study designs that sample fewer of these subjects will have little power to detect the association. An association with extreme obesity might well be masked by study design, and meta-analyses which disregard study design differences.

The tendency of higher OR observed in the general population-based studies (GP) and the obesity case-control studies compared to ‘healthy population’ (HP) studies could possibly reflect the association for extreme obesity compared to normal controls. We have classified studies as ‘ascertained for criteria related to a better health status’ (‘HP’) when patient groups were excluded or when the sample was ascertained based on working populations, which are known to be usually more healthy. We have performed this classification blinded for the study estimate to exclude bias from informative misclassification. It could be that the common rs7566605 directly or via tagging another possibly rare and quite penetrant variant does not do more much BMI throughout the distribution, but really puts participants into the very obese category. Thus an effect is picked up in the GP samples, but not in the HP studies with fewer extremely obese persons. This would also be in-line with (i) our more pronounced findings in the studies with a more recent BMI assessment might reflect this gene-environment interaction as well: assuming that a change in nutritional habits and physical activity contributed to the increase in obesity observed in the last decades, the studies with a more recent BMI assessment might reflect this more “modern” environment and the INSIG2-obesity association would emerge here more clearly.

Also, unknown epistatic interaction of the rs7566605 with one or other (rare) polymorphisms could lead to association with the more extreme obesity phenotype, with the INSIG2 gene being part of a complex that functions as a biological entity [SREBP, SCAP, INSIG2].

The importance of ascertainment of the study sample might be under-recognized so far. Monses and colleagues [16] have illustrated that ascertaining for or against disease would induce a bias in genetic association estimates when the genetic marker as well as the phenotype under study (here obesity) are associated with the disease. As obesity is associated with many chronic diseases such as type 2 diabetes and cardiovascular disease, exclusion of such study participants opens up for bias, if association of the SNP with the disease cannot be precluded. We had specifically excluded studies ascertained for disease and had also planned on separating HP from GP studies ahead of the analyses. We would like to highlight that we have adopted a very strict definition of GP and that there might be special advantages

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Table 2. Comparing more extreme degrees of obesity.

| BMI (kg/m²) | < 20 | 20–25 | 25–30 | 30–32.5 | 32.5–35 | 35–37.5 | 37.5–40 | > 40 | P-valueb |
|------------|------|--------|--------|---------|---------|---------|---------|-----|---------|
| All-CA     | %CC  | 10.4   | 10.7   | 10.5    | 10.5    | 11.9    | 11.2    | 12.0 | 12.5    | 0.0002 |
| N          | 66,213| 3314   | 22462  | 23072   | 5869    | 3578    | 2150   | 1404 | 3364    |
| All-CAa    | %CC  | 9.9    | 10.8   | 10.9    | 10.8    | 11.6    | 12.0    | 12.1  | 12.7    | 0.0008 |
| N          | 39,414| 2321   | 13671  | 3447    | 1962    | 1962    | 1162   | 829  | 2317    |
| GP         | %CC  | 10.7   | 10.8   | 10.2    | 10.4    | 12.0    | 12.4    | 12.6  | 13.5    | 0.007  |
| N          | 48,844| 2490   | 18551  | 18641   | 4300    | 2299    | 1111   | 620  | 832     |
| GPa        | %CC  | 10.6   | 10.8   | 10.8    | 10.7    | 12.01   | 12.9    | 12.1  | 13.7    | 0.005  |
| N          | 29,046| 1684   | 10975  | 10584   | 2546    | 1432    | 727    | 421  | 677     |
| HP         | %CC  | 10.6   | 10.4   | 11.9    | 9.7     | 8.1     | 7.5     | 13.0  | 8.6     | 0.345  |
| N          | 7640  | 273    | 2772   | 3288    | 679     | 323     | 147    | 77   | 81      |
| HPa        | %CC  | 8.3    | 10.1   | 11.8    | 10.1    | 7.0     | 9.6     | 17.6  | 2.9     | 0.758  |
| N          | 4911  | 121    | 1716   | 2279    | 457     | 186     | 83     | 34   | 35      |
| OB         | %CC  | 8.5    | 10.3   | 10.8    | 11.8    | 13.1    | 10.3    | 11.3  | 12.4    | 0.009  |
| N          | 9729  | 551    | 2139   | 1143    | 890     | 956     | 892    | 707  | 2451    |
| OBa        | %CC  | 8.1    | 11.9   | 10.0    | 12.4    | 12.2    | 10.8    | 11.5  | 12.4    | 0.054  |
| N          | 5457  | 516    | 980    | 842     | 444     | 344     | 352    | 374  | 1605    |

The data suggest an association when comparing more extreme degrees of obesity with normal controls, which is a potential explanation for the heterogeneity of the INSIG2 rs7566605 association with obesity (Hypothesis 2). Numbers stated are frequencies of risk genotype CC (C being the minor allele) across BMI categories for the Caucasian adult studies combined (All-CA) as well as stratified by study type (GP = general population, HP = healthy population, OB = obesity study). Also given are the p-values from testing for a trend of genotype frequencies across categories. (Not for All-NC or All-CH due to the few subjects in each category.)

Excluding studies published before the response letter by Herbert et al., December 2006 [1] in which the hypothesis of potential heterogeneity due to study design and a first call for this meta-analysis were stated (i.e., excluding American_Polish, NHS, KORA_S4, Essen_trios, EPIC_Norfolk, MRC_Ely, DESIR, SHIP, OB_adult, which are all Caucasian studies).

P-value testing for a trend across the BMI categories.

doi:10.1371/journal.pgen.1000694.t002
in using either disease-ascertained studies [28] or particularly healthy samples in other instances [29].

**Strengths and Limitations of This Study**

This meta-analysis has several strengths: (1) We have conducted a systematic approach by collecting all studies published before January 1, 2008, including seven studies that were unpublished at that time. (2) The meta-analysis is large including a total of 74,365 subjects. (3) We separated working tasks, with one researcher designing the analysis plan, recruiting studies, classifying studies by study type, and deciding upon compliance to inclusion criteria, while the other cared for the incoming data and performed the analysis. Therefore, design decisions were made in a blinded way, which guarded against subtle post-hoc data-driven analysis decisions, study selection bias, and informative misclassification of study design. (4) We collected data according to a strict protocol including standardized analysis from each study partner, with strong quality control of study-specific results. (5) We performed only a limited number of pre-defined subgroup analyses with some amendment during the review process. (6) We had a strong focus on the diversity of study design, which is unique in genetic epidemiological research at the time being and an issue probably under-recognized so far.

It might be considered a disadvantage that we did not include studies with subjects selected for diseases, particularly those associated with type 2 diabetes and thus a higher prevalence of obesity, as the association might be stronger in such studies. This might have been one reason for the initial investigation by Herbert and colleagues to detect this association, as mostly type 2 diabetes or asthma ascertained samples had been used. However, we excluded these samples by design in order to reduce heterogeneity or asthma. Excluding studies published before the response letter by Herbert et al., December 2006 [1], in which the hypothesis of potential heterogeneity due to study design and a first call for a meta-analysis were stated (i.e. excluding American_Polish, NHS, KORA-S4, Essen_trios, EPIC_Norfolk, MRC_Ely, DESIR, SHIP, OB_adult).

**Hypotheses 3–5**

Finally, it might be considered a disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motivated by the early published studies, which are included in this meta-analysis; to further disadvantage that our hypotheses were motiva...
disadvantage that we were not able to recruit enough non-Caucasian adult or children studies for a conclusive comparison with the Caucasian adult studies.

Conclusions

This pooled analysis including all study designs does not provide evidence for overall association of the INSIG2 rs7566605 CC genotype with increased risk of obesity compared to the CG or GG genotypes. Our data suggest an association with extreme degrees of obesity and consequently heterogeneous effects from different study designs may mask an underlying association when unaccounted for. The importance of study design might be underestimated in gene discovery and association replication so far.

Materials and Methods

Meta-Analysis Concept

We designed our meta-analysis as a pooled analysis of study-specific association estimates according to a standardized protocol (see ‘data form’, Text S1B, and pre-defined analysis plan, Text S1C) with an amendment added during the review process (Text S1D).

Our eligibility criteria for studies were (i) data available on BMI, the INSIG2 rs7566605 SNP genotypes, age and sex, (ii) sample size of at least 200 subjects, (iii) ethical approval, and (iv) either general population-based (GP), ascertained for reasons related to a better health status such as studies including only subjects in the workforce or studies excluding subjects with diseases (‘healthy population’, HP), or designed specifically to study obesity such as obesity case-control or obesity family studies (OB). We excluded all studies selecting subjects for any disease. More information on the classification of studies by study type, see Text S1E. We did not exclude on any age or ethnicity criteria to allow exploration of potential heterogeneity. We controlled for study selection bias by separating the two main tasks between the two first authors: IMH took care of study recruitment, compliance to inclusion criteria, and classification of studies by study type, and CH performed quality control and statistical analysis.

Study Recruitment, Collection of Aggregated Data, and Quality Control

We identified all eligible studies published before January 1, 2008 by a systematic PubMed literature search using the search terms ‘INSIG2’ OR ‘INSIG-2’ OR ‘rs7566605’. Additionally we identified unpublished studies through contacting researchers in the field by making a call for this meta-analysis in several consortia (GIANT, KORA-500K, IL-6-consortium), in the letter to Science by Herbert and colleagues [1], in the paper by Lyon and colleagues [12], and in meeting presentations. We sent out and collected standardized data forms, and verified all entries for within-plausibility as well as consistency with publications, if available. We made plausibility checks by use of double within-plausibility as well as consistency with publications, if available. We made plausibility checks by use of double

Statistical Analysis

For each study, OR estimates comparing the odds of obesity (BMI≥30 kg/m²) for subjects with the minor-allele homozygous genotype (CC) with subjects of the other genotypes combined (CG, GG), thus assuming a recessive model, were calculated using logistic regression adjusting for age and sex. We also collected OR estimates with standard errors (SE) for the odds of more extreme degrees of obesity (i.e. subjects with BMI≥32.5, 35.0, 37.5, or 40.0 kg/m²) compared to various degrees of leanness (i.e. subjects with BMI<30, 25, or 20 kg/m²). Furthermore, we collected summary statistics (mean and SE) on the difference in mean BMI between subjects with the CC genotype compared to subjects with the CG or GG genotypes using linear regression adjusted for age and sex. For each study, analyses stratified for sex or age (with a cut-off at 50 years) were performed as well. Among the six studies from non-Caucasian populations, two studies had too few (<3) subjects among the obese with the CC genotype to be included into the dichotomous obesity analysis, while they were included for the quantitative BMI analysis.

For the meta-analysis, we combined beta-estimates among Caucasian adult studies (All-CA) followed by a stratified analysis by study type (GP, HP, OB) and combined estimates among non-Caucasian (All-NC) or children studies (All-CH), see ‘amendment to analysis plan’, Text S1D.

The following was only performed on Caucasian adults as the number of available non-Caucasians or children was too low. We tested for differential association between the GP, HP, or OB studies applying a t-test on the combined beta-estimates and correcting p-values for testing three subgroups. We tested for a trend in CC genotype frequencies across the different BMI categories and tested for differential associations separating the studies for higher or lower obesity prevalence, higher or lower mean age of study subjects, or for a more or less recent BMI assessment using a t-test on the combined beta-estimates. This was complemented by sensitivity analyses stratifying on sex, age, publication status, and type of BMI assessment. As the hypotheses were motivated by the early published studies mentioned in the letter to Science by Herbert and colleagues [1], we repeated all analyses with exclusion of these hypotheses-generating studies.

In all analyses, between-study heterogeneity was tested by the $x^2$-based Q-statistic and quantified by I² as a measure of the proportion of variance between the study-specific estimates that is attributable to between-study difference rather than random variation. We pooled study-specific estimates according to the inverse-variance weighted fixed effect or the DerSimonian and Laird random effects model [30]. Heterogeneity was considered to be significant at the 10% level. All statistical analyses were performed with SAS (statistical analysis software, SAS institute, Inc.). Forest plots were prepared using Review Manager software (Cochrane Collaboration, Copenhagen, DK).

Supporting Information

Figure S1 Forest plot on estimates of the INSIG2 rs7566605 association with BMI. Association of the INSIG2 SNP with BMI for (A) the Caucasian adult general-population based studies (GP), (B) the Caucasian adult studies selected for healthy population (HP), (C) the Non-Caucasian adult studies, (D) the pediatric studies. Shown are recessive model beta estimates comparing the CC genotype versus CG or GG for each study and pooled estimates. The fixed effect (FE) model beta is shown in case of no significant heterogeneity as tested by the Q-statistics; a random effect (RE) model is shown otherwise. Found at: doi:10.1371/journal.pgen.1000694.s001 (0.44 MB TIF)

Table S1 Characterization of eligible and recruited studies. Found at: doi:10.1371/journal.pgen.1000694.s002 (0.12 MB DOC)
Table S2 More details on study characteristics. Found at: doi:10.1371/journal.pgen.1000694.s003 (0.20 MB DOC)

Table S3 Stronger genetic effects when comparing more extreme degrees of obesity as an explanation of the heterogeneity of the INSIG2 rs7566605 association (Hypothesis 2). Pooled association estimates for increasing degrees of obesity for all Caucasian adult studies combined (All-CA) as well as stratified by study type (GP = general population, HP = healthy population, OB = obesity study), for all Non-Caucasian studies (All-NC), and for children studies due to non-comparability of BMI categories. Numbers stated are ORs comparing cases versus controls (p-values) from the pooled analysis (fixed and random effects), number of cases/controls (number of pooled studies), and the I² (p-value from Q-statistics). Found at: doi:10.1371/journal.pgen.1000694.s004 (0.07 MB DOC)

Table S4 Sensitivity analyses for the association of the INSIG2 SNP with obesity regarding sex, age, published or unpublished status of studies, self-reported or measured BMI. Stated values are ORs (p-values) based on fixed or random effects models, the I² (p-value of Q test) for each group, and p-values testing for difference between the fixed effect [random effect] ORs of the two corresponding groups. (Not for HP, OB, All-NC, or All-CH due to low numbers of studies.) Found at: doi:10.1371/journal.pgen.1000694.s005 (0.07 MB DOC)

Table S5 Main results of pooled association of the INSIG2 SNP with body-mass-index (BMI). The analyses for all Caucasian adult studies combined (All-CA) as well as stratified by study type (GP = general population, HP = healthy population, OB = obesity study), for all Non-Caucasian studies (All-NC), and for the children studies (All-CH) indicated some difference between GP and HP studies (Hypothesis 1). Numbers stated are recessive model beta-estimates (p-values), i.e., mean difference of BMI between subjects with the CC genotype compared to subjects with the CG or GG genotype, using fixed or random effects models, the I² (p-value of Q-statistics), and p-values testing for pair-wise difference between GP, HP, or OB studies. Found at: doi:10.1371/journal.pgen.1000694.s006 (0.04 MB DOC)

Table S6 Exploring potential sources of heterogeneity of the INSIG2 rs7566605 association with BMI (Hypotheses 3-5). Stated values are pooled beta-estimates (p-values) based on fixed or random effects model, I² (p-values of Q-test) for each group, and p-values testing for difference between the beta-estimates of the two corresponding groups. (Not for HP, NC, or CH due to low numbers of studies.) Found at: doi:10.1371/journal.pgen.1000694.s007 (0.08 MB DOC)

Table S7 Sensitivity analyses for the association of the INSIG2 SNP with BMI regarding sex, age, published status, or self-reported or measured BMI. Stated values are beta-estimates (p-values) based on fixed and random effects models, the I² (p-value of Q test) for each group, and p-values testing for difference between the beta-estimates of the two groups. (Not for HP, All-NC, or All-CH due to low numbers of studies.) Found at: doi:10.1371/journal.pgen.1000694.s008 (0.07 MB DOC)

Text S1 References of included published studies; data form sent to each study partner; predefined analysis plan; amendment to analysis plan; and classification of studies due by study type. Found at: doi:10.1371/journal.pgen.1000694.s009 (0.18 MB DOC)

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Author Contributions

Conceived and designed the experiments: IMH CH. Performed the experiments: IMH CH. Analyzed the data: IMH CH RJFL VA SAA KA HB PB HB CB MC JAC DC CD JCE EF FF PF JH RAH AR MH FBH [JAH SEH SCH TI MR] MK BK HK JK LAL BL SL HNL DM KLM VM AN TTN BP LP LQ TR DR SS SS AS GT UT HV KSV NJW DMW SY CL MM CL JNH NL HEW. Contributed reagents/materials/analysis tools: IMH CH RJFL FK VA SAA KA HB PB HB CB MC JAC DC CD JCE EF FF PF JH RAH AR MH FBH [JAH SEH SCH TI MR] MK BK HK JK LAL BL SL HNL DM KLM VM AN TTN BP LP LQ TR DR SS SS AS GT UT HV KSV NJW DMW SY CL MM CL JNH NL HEW. Wrote the paper: IMH CH AL NL. Advised with paper writing: RJFL. Advised in design and paper writing: FK. Consulted in the revision: AL. Advised in study design: CL. Motivated the project and consulted in study design: HEW.

References

1. Herbert A, Gerry NP, McQueen MB, Heid IM, Pfeuffer A, et al. (2006) A common genetic variant is associated with adult and childhood obesity. Science 312: 279-283.
2. Yabe D, Brown MS, Goldstein JL (2002) Insig-2, a second endoplasmic reticulum protein that binds SCAP and blocks export of sterol regulatory element-binding proteins. Proc Natl Acad Sci U S A 99: 12753–12758.
3. Ciudal M, Nutile T, Dalmasso C, Sorice R, Bellenguez C, et al. (2008) Identification and replication of a novel obesity locus on chromosome 1q21 in isolated populations of Cilento. Diabetes 57: 783-789.
4. Kumar J, Sunikshala RR, Karthikeyan G, Sengupta S (2007) The common genetic variant upstream of INSIG2 gene is not associated with obesity in Indian population. Clin Genet 71: 415–418.
5. Heid IM, Vollmert C, Hinney A, Doering A, Geller F, et al. (2005) Association of rs7566605 with Obesity. Meta-Analysis on the INSIG2 rs7566605 with Obesity (Hypothesis 1). Numbers stated are ORs based on fixed or random effects model, I² (p-values of Q-test) for each group, and p-values testing for difference between the beta-estimates of the two corresponding groups. (Not for HP, NC, or CH due to low numbers of studies.) Found at: doi:10.1371/journal.pgen.1000694.s008 (0.18 MB DOC)
15. Hill JO, Peters JC (1998) Environmental contributions to the obesity epidemic. Science 280: 1371–1374.
16. Monsees GM, Tamimi RMK, Kraft P (2009) Genome-wide association scans for secondary traits using case-control samples. Genet Epidemiol;epub Apr 13.
17. Nebeil A, Crouscher PJ, Stiegeler R, Nikolaus S, Krawczak M, et al. (2005) No association between microsomal triglyceride transfer protein (MTP) haplotype and longevity in humans. Proc Natl Acad Sci U S A 102: 7906–7909.
18. Marvelle AF, Lange LA, Qin L, Adair LS, Moblie KE (2008) Association of FTO with obesity-related traits in the Cebu Longitudinal Health and Nutrition Survey (CLHNS) cohort. Diabetes 57: 1987–1991.
19. Loefgen E, Pouta A, Wendt LV, Tapanainen J, Isoaervi JI, et al. (2008) Epilepsy in the northern Finnish birth cohort 1966 with special reference to fertility. Epilepsy Behav 14: 102–107.
20. Jaervelin MR, Sovio U, Kind V, Lauren L, Xu B, et al. (2004) Early Life Factor and Blood Pressure at Age 51 Years n the 1966 Northern Finland Birth Cohort. Hypertension 44: 836–846.
21. Andreasen CH, Mogensen MS, Borch-Johnsen K, Sandbaek A, Lauritzen T, et al. (2008) Non-replication of genome-wide based associations between common variants in INSIG2 and PFKP and Obesity Studies of 18,014 Danes. PloS ONE 3: 2072. doi:10.1371/journal.pone.0002072.
22. Liu YJ, Liu XG, Wang L, Dina C, Yan H, et al. (2008) Genome-wide association scans identified CTNNBL1 as a novel gene for obesity. Hum Mol Genet 17: 1803–1813.
23. Zhao Y, Chan MY, Zhou S, Heng CK (2008) Effects of atherogenic diet and atorvastatin treatment on gene expression profiles in the C57BL/6J mouse liver. Gene Expr 14: 149–158.
24. Koch A, Konig B, Spielmann J, Leitner A, Stangl GI, et al. (2007) Thermally oxidized oil increases the expression of insulin-induced genes and inhibits activation of sterol regulatory element-binding protein-2 in rat liver. J Nutr 137: 2018–2023.
25. Reinehr T, Hinney A, Nguyen TT, Hebebrand J (2008) Evidence of an influence of a polymorphism near the INSIG2 on weight loss during a lifestyle intervention in obese children and adolescents. Diabetes 57: 625–626.
26. Reinehr T, Hinney A, Toschke AM, Hebebrand J (2009) Aggravating effect of INSIG2 and FTO on overweight reduction in a one-year lifestyle intervention. Arch Dis Child Pub ahead of print.
27. Orkunoglu-Suer FE, Gordish-Dressman H, Clarkson PM, Thompson PD, Angelopoulos TJ, et al. (2008) INSIG2 gene polymorphism is associated with increased subcutaneous fat in women and poor response to resistance training in men. BMC Med Genet 9: 117.
28. Hinney A, Nguyen TT, Scherag A, Friedel S, Bronner G, et al. (2007) Genome wide association (GWA) study for early onset extreme obesity supports the role of fat mass and obesity associated gene (FTO) variants. PLoS ONE 2: e1361. doi:10.1371/journal.pone.0001361.
29. Heid IM, Wagner SA, Gohike H, Isgheer B, Mueller JC, et al. (2006) Genetic architecture of the APM1 gene and its influence on adiponectin plasma levels and parameters of the metabolic syndrome in 1,727 healthy caucasians. Diabetes 55: 375–384.
30. DerSimonian R, Laird N (1996) Meta-analysis in clinical trials. Control Clin Trials 7: 177–188.