On the relative ages of the $\alpha$-rich and $\alpha$-poor stellar populations in the Galactic halo

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ABSTRACT

We study the ages of $\alpha$-rich and $\alpha$-poor stars in the halo using a sample of F and G dwarfs from the Sloan Digital Sky Survey (SDSS). To separate stars based on [$\alpha$/Fe], we have developed a new semi-empirical spectral index-based method and applied it to the low-resolution, moderate signal-to-noise SDSS spectra. The method can be used to estimate the [$\alpha$/Fe] directly providing a new and widely applicable way to estimate [$\alpha$/Fe] from low-resolution spectra. We measured the main-sequence turnoff temperature and combined it with the metallicities and a set of isochrones to estimate the age of the $\alpha$-rich and $\alpha$-poor populations in our sample. We found all stars appear to be older than 8 Gyr confirming the idea that the Galactic halo was formed very early on. A bifurcation appears in the age–metallicity relation such that in the low-metallicity regime the $\alpha$-rich and $\alpha$-poor populations are coeval while in the high-metallicity regime the $\alpha$-rich population is older than the $\alpha$-poor population. Our results indicate the $\alpha$-rich halo population, which has shallow age–metallicity relation, was formed in a rapid event with high star formation, while the $\alpha$-poor stars were formed in an environment with a slower chemical evolution time-scale.

Key words: Galaxy: halo.

1 INTRODUCTION

We live in a typical spiral galaxy in the Universe, but what makes the Milky Way special is that only in our Galaxy can we probe the chemical, kinematical and spatial distribution of individual stars with high accuracy. Different patterns in these distributions are used to define different stellar populations and studying these in detail is crucial to understand how spiral galaxies form and evolve. In this context, the stellar halo is of paramount importance to reveal the early stages of the Milky Way because these stars are mostly very old and metal-poor (e.g. Unavane, Wyse & Gilmore 1996; McWilliam 1997; Helmi 2008; Jofré & Weiss 2011; Kalirai 2012).

After the classical work of Eggen, Lynden-Bell & Sandage (1962), where the stellar halo is composed of a simple stellar population which formed from the monolithic collapse of the protogalactic cloud, many works have shown that the stellar halo is rather complex with components of stars being accreted after the initial collapse (Searle & Zinn 1978; Ibata, Gilmore & Irwin 1994; Belokurov et al. 2006). Much of the discussion in recent years has been focused on ways to quantify the amount of different stellar populations in a way to ascertain the importance of the merging history in the formation of the Milky Way. The works of Carollo et al. (2007, 2010), for example, have shown observational evidence that there is an outer stellar halo, which is dominant at galactocentric distances larger than 15 kpc, mostly formed from accreted subunits and an inner stellar halo, confined to galactocentric distances less than 15 kpc, which seems to be rather smooth, with no predominant subunit like the outer halo. However, these results are still debated (see discussion in Schönrich, Asplund & Casagrande 2011; Ivezić, Beers & Juric 2012).

Interestingly, inner halo stars seem to have two chemical patterns: a classical one with [$\alpha$/Fe] $\sim$ +0.40 which is related to the product of star formation in the large initial collapsing protogalactic gas cloud (often thought of as the in situ population); and another one with [$\alpha$/Fe] $\sim$ +0.20, which is related to the formation of stars in an environment of lower star formation rate, typically in smaller gas regions (e.g. Nissen & Schuster 2010, hereafter N10). The latter are attributed to have extragalactic origins that were accreted on to the Milky Way after the formation of the main population (i.e. accreted stars). These two populations are commonly referred as ‘$\alpha$-rich’ and ‘$\alpha$-poor’ (N10), which have been subject of interest among other studies (Ramírez, Meléndez & Chanamé 2012; Sheffield et al. 2012). These two populations have been extensively studied in a series of papers by Schuster & Nissen. They have used a sample of high resolution and high signal-to-noise ratio (SNR) spectra of halo stars to show that these two populations are distinct in kinematics and abundances of $\alpha$-elements (N10), but indistinguishable from other chemical abundances such as Li and Mn (Nissen & Schuster 2011, 2012). One interesting question arises with the findings of
N10: do the $\alpha$-poor and $\alpha$-rich populations have the same age and/or is there correlation between age and metallicity? Determining the age differences between these two populations will help distinguish between the formation and assembly time-scales of the Galactic halo.

Schuster et al. (2012) attempted to answer this question by finding that the $\alpha$-poor stars in their sample are $\sim$2–3 Gyr younger than the $\alpha$-rich population. This is in favour of the models of Zolotov et al. (2009, 2010). Yet, the recent simulations of Font et al. (2011) found that in situ stars can be as much as 3–4 Gyr younger than the accreted population. Schuster et al. (2012) argued that their $\alpha$-rich and $\alpha$-poor populations can be explained by a scenario where an initial disc/bulge formed in a monolithic collapse producing the $\alpha$-rich population. At later times, the $\alpha$-rich stars in the primeval disc were scattered into the halo via merging events that subsequently populated the $\alpha$-poor component of the halo. However, the conclusions of Schuster et al. (2012) were drawn from only nine stars lying in the metallicity range of $-0.4 < [\text{Fe/H}] < -1.40$ dex. This made us wonder whether a difference in age between $\alpha$-poor and $\alpha$-rich can also be found using the data from the Sloan Digital Sky Survey (SDSS; York et al. 2000), which contains thousands of halo stars extending the metallicity domain towards much lower metallicities compared to Schuster et al. (2012). Although the spectra from SDSS have much lower resolution, it is still possible to rank the metal-poor stars in $\alpha$-abundance space.

Most of the current methods developed for measuring $\alpha$-abundances with low-resolution spectra attempt synthetic spectral matching (e.g. Lee et al. 2011). In these methods, a grid of synthetic spectra is constructed with relatively fine spacing in $[\alpha/\text{Fe}]$ and degraded to low resolution. Other methods (e.g. Franchini et al. 2010, 2011), use a set of Lick indices to measure the $[\alpha/\text{Fe}]$. Franchini et al. (2011) determined the $[\alpha/\text{Fe}]$ of a sample of F, G and K stars observed with SDSS within $\pm0.04$ dex (accounting for the internal errors only). These methods heavily rely on having a good representation of real spectra through a synthetic grid of spectra and fairly well-determined stellar parameters. The primary disadvantage to those methods, is that the estimated $[\alpha/\text{Fe}]$ is strongly affected by the stellar parameter uncertainties because of the degeneracies that exist between the stellar parameters and the $[\alpha/\text{Fe}]$. These factors make estimating $[\alpha/\text{Fe}]$ at low metallicity extremely difficult. We developed a new method which moves beyond the grid matching techniques, of current methods by defining an index, with the aid of a synthetic grid of spectra, which is computed using only the observed spectra. In this way, our method is a semi-empirical way to estimate $[\alpha/\text{Fe}]$. Using this method, we can rank our sample based on their $\alpha$-abundance in a more model-independent way than the current methods. With this method, we aim to measure the age-metallicity relation of $\alpha$-rich and $\alpha$-poor halo field stars separately.

Due to poor distance estimates, we did not attempt to determine ages with the standard isochrone fitting technique (for a discussion on the method consult, Soderblom 2010) which was also employed by Schuster et al. (2012). The large number of stars in our sample allows us to group stars in metallicity and determine the age of that population very accurately if there is a coeval dominant population (Unavane et al. 1996; Schuster et al. 2006; Jofrè & Weiss 2011). This way of determining ages relies on using the colour (or temperature) of the main-sequence turnoff of the population, which presents a sharp edge in the temperature distribution. By comparing the turnoff temperature of an $\alpha$-rich population at a certain metallicity and an $\alpha$-poor population of that metallicity, we can quantify the age difference between those populations. That age difference (if any), using a much larger sample of stars and a larger metallicity coverage, provides clues as to when the accreted stars of the inner halo formed relative to the in situ ones. That information is valuable for constraining theoretical models of the Milky Way formation.

This paper is organized in the following way. In Section 2, we define the sample of SDSS halo stars for which we estimated the $\alpha$-abundances and ages. In Section 3, we describe our new method to categorize low-resolution spectra based on their $\alpha$-abundances in the regime of the Galactic halo. In Section 4, we validate our method. In Section 5, we employ the method to split our sample into a $\alpha$-rich and $\alpha$-poor population for which we determine the ages, and their errors in each population. In Section 6, we discuss the results and their implications for the formation of the Galactic halo. Finally, we summarize our findings in Section 7.

2 DATA

This study made use of the SDSS/DR9 (Ahn et al. 2012) and the SEGUE data base (Yanny et al. 2009). SEGUE/SDSS provides approximately half a million low-resolution ($R \sim 2000$) spectra which have stellar parameters estimates from the SEGUE stellar parameter pipeline (SSP; Lee et al. 2008a). The spectra have wavelength coverage of 3900–9000 Å. We are interested in halo F- and G-type dwarfs, which allows us to determine the location of the turnoff. We selected these dwarf stars via a colour cut requiring the reddened $g-r$ colours to be in the range: $0.1 < (g-r) < 0.4$ mag. We further required the SNR achieved to be at least 40 (see Section 3.6). To maximize the number of halo stars while deseleting other Galactic components we focused on the metallicity below $-0.80$ dex and further required the absolute Galactic latitude, $|b|$, to be larger than 30°. While we expect some contamination from the thick disc in the most metal-rich bin, given the metallicity distribution function of Kordopatis et al. (2013), it is likely that our sample is dominated by halo stars. However, accurate distances and proper motions would be needed to fully resolve the space motion of the stars to study the contamination fraction.

We made additional cuts on the adopted SSPP parameters such that they are within the stellar parameter range defined in Section 3.4. Finally, we co-added any duplicated spectra to increase the SNR. Our final sample contains 14 757 unique objects. A colour–colour diagram of the sample can be found in Fig. 1. The $(g-r)_0$ colour is directly related to the temperature of the star. Thus, Fig. 1 illustrates the main-sequence F and G dwarf stars of our sample are plentiful until $(g-r)_0 \sim 0.2$ mag which is approximate location of the turnoff. Above this temperature, i.e. smaller $(g-r)_0$ colours, are likely blue stragglers. Given the SNR cut and the fact that we select dwarfs; our sample is limited to stars near the sun (likely within 1–2 kpc) and thus the inner halo.

3 A METHOD TO ESTIMATE $\alpha$-ABUNDANCES

3.1 Grid of synthetic spectra

We used a grid of synthetic spectra to develop and test our spectral-index method (SIM). In the grid, $\alpha$-enhancement is achieved by increasing (or decreasing) in lockstep the individual $\alpha$-elements (Ca, Ti, Si, Mg, O) from their solar values. The synthetic spectra make use of the 1D LTE MARCS model atmospheres of Gustafsson et al. (2008) which have a variety of $\alpha$-abundances. The synthetic grid was created using the TURBOSPECTRUM synthesis code (Alvarez & Plez 1998; Plez 2012) which uses the line broadening treatment described by Barklem & O'Mara (1998). Solar abundances were
taken from Asplund, Grevesse & Sauval (2005). Atomic lines used by TURBOSPECTRUM are sourced from VALD, Kupka & Ryabchikova (1999), Hill et al. (2002) and Masseron (2006). Line-lists for the molecular species are provided for CH (Masseron et al. 2014), and CN, NH, OH, MgH and C2 (Masseron, in preparation); the lines of SiH molecules are adopted from the Kurucz line-lists and those from TiO, ZrO, FeH, CaH from Plez (private communication). Microturbulence velocity for each spectrum were estimated using a polynomial relationship (Bergemann et al., in preparation) between microturbulence velocity and surface gravity developed for the Gaia-ESO Survey (Gilmore et al. 2012). The final grid covers 3000 K ≤ Teff ≤ 8000 K in steps of 200 K, 0.0 ≤ log g ≤ 5.0 in steps of 0.2 dex, and −3.0 ≤ [Fe/H] ≤ +1.0 in steps of 0.1 dex and −0.1 ≤ [α/Fe] ≤ +0.4 dex in steps of 0.1 dex. The synthetic grid has only been used to provide a starting point to inform our placement of spectral bands which are sensitive to [α/Fe].

3.2 Spectral Index Method

Our aim is to classify stars as α-rich or α-poor on the basis of an index that is relatively insensitive to stellar parameters. Following similar methods, often employed when studying stellar populations in other galaxies (e.g. Thomas, Maraston & Bender 2003), we used a spectral index that is very sensitive to α-abundances. This allows us to use a simplistic approach to find the α-abundances using low-resolution SDSS spectra. The index was built with a combination of (1) spectral bands that are sensitive to α-abundances and (2) control bands whose response to the stellar parameters mimic the response function of the α-sensitive bands. All spectral bands must be large enough as to not be dominated by noise, yet not too large to be dominated by broad structures in the spectrum. Since we are interested primarily in [α/Fe], we only choose bands which are sensitive to α-abundances to be the main driver of our index diagnostic. We employed a semi-automatic method to explore the synthetic spectral grid to find moderate-size (15 Å < band width < 80 Å) spectral bands. We found five bands that are sensitive to the α-abundances (listed in the top part of Table 1). We combined these bands in a linear way to obtain an index sensitive to [α/Fe].

The control bands, on the other hand, were found via an iterative process by systematically scanning a grid of possible control bands. A good control band is one that minimizes the scatter in the index over a range of stellar parameters while maximizing the mean difference between the index at [α/Fe] = 0.0 dex and [α/Fe] = +0.4 dex. We found three control bands, whose wavelength ranges are listed at the bottom of Table 1. Fig. 2 shows two synthetic spectra, one of with solar-scaled α-abundances (black solid line) and the other that is α-enhanced (green dotted line). The spectral regions are designated with red dashed lines corresponding to control bands and blue solid lines for α-sensitive bands.

With the set of spectral bands defined in Table 1, we define our α index diagnostic by

\[
\text{Index} = \frac{5\text{Mgb} + 4\text{KTi} + 3\text{KSi} + \text{KCa} + \text{K}1}{\text{CB1} + \text{CB2} + \text{CB3}}.
\]

The weights in the equation are motivated by the line-strengths of the α-sensitive bands, and were determined after trying several combinations. The proposed method differs from current methods which rely on a synthetic grid to model out the effects of the stellar parameters (e.g. Thomas et al. 2003; Franchini et al. 2010, 2011). The advantage of our method is that we can use the spectrum alone to estimate the [α/Fe] with relatively little effect from the uncertainties in the stellar parameters provided they are within the stellar parameter range defined in Section 3.4 below.

From a physical point of view, the numerator of the index is defined by spectral bands that are centred near the spectral features of the α-elements. Therefore, an increase in the α-abundance causes an increase in the strength of the α-sensitive spectral bands. An example of this can be seen in the KSi spectral band in Fig. 2 which is centred on a series of Ti and Si features. The control bands (e.g. CB1) in the denominator are centred near Fe-peak elements (mostly Fe spectral features) yet have very similar transition properties (e.g. oscillator strength and excitation potential) as the α-sensitive bands. As a result of this, the index tracks α-abundances divided by Fe-peak abundances while simultaneously controlling for the stellar parameters. We note that the CB2 and KCa bands overlap. While the Fe and Cr lines in the CB2 band overlap in wavelength space with the KCa band, the response of the CB2 band to the stellar parameters is different than the α-sensitive lines in the KCa band. Tests have shown that excluding the overlapping KCa band region from the CB2 band increases the scatter in the index at high α-abundance by a factor of 1.4 leading to a less precise estimate.

3.3 Processing spectra

Synthetic and object spectra were pre-processed before searching for and measuring spectral features. All spectra were pre-processed in the same way so that it is possible to compare them directly. We applied this method to the SDSS spectra and as such their spectral parameters (e.g. resolution, etc.) defined the pre-processing.
Figure 2. Illustration of the KSi and Mgb (blue shaded regions), CB1 (red shaded region) spectral bands for an \(\alpha\)-rich (green dotted) and \(\alpha\)-poor (black solid) spectra with \(T_{\text{eff}} = 6200\) K, \(\log g = 4.4\) dex, and \([\text{M/H}] = -1.50\) dex.

parameters. We started by converting the wavelength scale of all spectra to an air-based system. Synthetic, and validation spectra were degraded to a resolution of 2000 using a Gaussian kernel. All spectra were resampled to 0.5 Å pixels over a wavelength range of 4000–7000 Å.

We defined the local continuum via a boosted-median pseudo-continuum (BMC; Rogers et al. 2010). This method is defined by taking a window around any particular wavelength and setting the continuum to be at a prescribed percentile of the fluxes within the window. The BMC method is based on two parameters: (1) the width of the window around a given wavelength and (2) the percentile of the flux which defines the pseudo-continuum. We started by assuming the parameters used in Rogers et al. (2010), a window of 100 Å with a percentile of 90 per cent. The effect of the choice of these parameters is analysed in detail in Section 3.7.

3.4 Performance of the index on synthetic spectra

We are interested primarily in F, G and K dwarf stars and thus we restricted the stellar parameters to:

(i) \(5000\) K \(\leq T_{\text{eff}} \leq 7000\) K 
(ii) \(\log g \geq 3.5\) dex 
(iii) \(-2.0\) dex \(\leq [\text{M/H}] \leq 0.0\) dex.

Selecting warm stars means that we avoided significant molecular features that would require a more careful calibration in the index. Therefore, we required that the effective temperature must be at least 5000 K. However, if the temperature is too high, the lines we need for our diagnostic will be weaker and thus we set an upper limit in temperature at 7000 K, potentially excluding any very young metal-poor stars. We required the metallicities to be above \(-2.0\) dex because below that limit it can be difficult, especially at high temperatures, to estimate the \(\alpha\)-abundance given the lack of strong lines (for more discussion see Section 3.5). The \(\log g\) cut is to ensure we have selected dwarf stars. It is interesting to note that this simple method is used to search for spectral peculiarities, in our case stars with strong features around the \(\alpha\)-elements. However, the method could, in theory, be recalibrated to search for other scientifically exciting targets such as \(s\)-process or \(r\)-process enriched stars or others. This method could also be adapted to determine the abundance of other chemical species using low-resolution spectra.

Figure 3. The index value as a function of [\(\alpha/\text{Fe}\)] colour coded by \(T_{\text{eff}}\) (top panel), [Fe/H] (middle panel), \(\log g\) (bottom panel) for a sample of 200 synthetic spectra.

In Fig. 3, we plot the value of the index for a sample of 200 synthetic stars (selected from the parameter space and synthetic grid above), where the \(y\)-axis labels the index value and the \(x\)-axis is the synthetic value for [\(\alpha/\text{Fe}\)]. The colour code represents the temperature, metallicity and surface gravity in the top, middle and lower panels, respectively. We see a tight, linear correlation between the index value and the [\(\alpha/\text{Fe}\)], such that high index values indicate high \(\alpha\)-abundances. There does not appear to be any major systematic variations as a result of the stellar parameters. Fig. 3 shows that it is possible to distinguish between an \(\alpha\)-rich and \(\alpha\)-poor population as the mean index of stars with [\(\alpha/\text{Fe}\)] \(\leq +0.25\) dex is smaller than the mean index of stars with [\(\alpha/\text{Fe}\)] \(> +0.25\) dex. A star with an index below approximately 7.4 can be interpreted
as $\alpha$-poor while a star with an index about 7.6 $\alpha$-rich. Fig. 3 also suggests that it may be possible to calibrate the index in a linear way to estimate the [$\alpha$/Fe] directly (which we explore in Section 4.4).

### 3.5 Effects of stellar parameters

In this section, we study the effect of the stellar parameters on the index. The index is a linear combination of $\alpha$-sensitive spectral bands divided by a set of control spectral bands. Ideally, the index should be constant with a given [$\alpha$/Fe] regardless of the stellar parameters. In practice, the stellar parameters have a small effect on the index. To study this effect, we computed the index over a broad range of stellar parameters (outlined in Section 3.4) at a constant [$\alpha$/Fe]. Fig. 4 shows the value of the index (colour) as a function of both temperature and metallicity. The plot illustrates that the index is stable and only varies of the order of 1–2 per cent at effective temperatures below ~6500 K. However, the index seems to vary significantly at the very metal-poor and high temperature end (bottom-right corner of Fig. 4). This is expected as the metal-poor hot stars will likely have significantly weaker lines. From this figure, we can see that the method does not preform nearly as well for the hot (particularly at temperatures above 7000 K), metal-poor stars.

### 3.6 Effects of Signal-to-Noise

To study the effect of both the SNR and the stellar parameters on the stability of the index, we have completed noise-injection experiments. We used the synthetic grid of spectra described in Section 3.1. Once these spectra were pre-processed, we added Gaussian white noise to simulate spectra at SNR between 20 and 100 typical for SDSS data. We then preformed a BMC continuum normalization. Finally, we calculated the index and plotted the variation in the index as a result of the stellar parameters at each SNR for [$\alpha$/Fe] = 0.00 dex and [$\alpha$/Fe] = +0.40 dex. Both the SNR and stellar parameters cause a variation in the index at a constant [$\alpha$/Fe].

Fig. 5 shows that it is possible to separate [$\alpha$/Fe] = 0.00 and [$\alpha$/Fe] = +0.40 dex at SNR larger than 40. We find, at an SNR = 40, the total internal error on the index, which is defined as its mean variation, is found to be 0.60. Fig. 4 illustrates that this value can be decreased significantly by further restricting the parameter space below 6500 K. For example, by restricting the parameter space to metal-poor halo turnoff stars ($6000 < T_{\text{eff}} < 6500$ K, log $g > 3.5$ dex and $-2.0 < [\text{Fe/H}] < -0.80$ dex), the total internal error 0.20 instead of 0.60.

### 3.7 Effect of pseudo-continuum placement

We used the BMC method to determine the local continuum for our spectra that has the benefit of better control the continuum in the local region of our bands. It also has the advantage of reducing the uncertainty in the total flux of our measured spectral bands (Rogers et al. 2010). The BMC method is sensitive to two parameters and a balancing act must be played with both. These parameters are as follows.

(i) The width of window around any spectral band – the width must be large in order to avoid small-scale fluctuations and noise in the spectra but not too large as to encode large-scale structure (e.g. G band at ~4300 Å).

(ii) The percentile of flux within the window which is defined as the continuum – the percentile must be high enough to ‘see’ the true continuum but cannot be so high that it becomes only sensitive to noise.

Therefore, it is useful to study the sensitivity of our index to these two parameters that define the placement of the continuum. To do this, we took a single random synthetic spectrum and plotted the contours of the measured index for that spectrum as a function of both the width and percentile. The index should be insensitive to our choice of BMC parameters and thus we searched for regions where the derivative of the index to the BMC parameters is minimized. Remembering that the density of contours is proportional to the derivative of the index as a function of the BMC parameters, we selected a region that the index is very stable over a range of BMC parameters. The choice of BMC parameters adopted from Rogers et al. (2010), which is the ‘X’ in Fig. 6, worked very well because the index does not seem to depend strongly on the choice of BMC parameters. This experiment has shown that as long as the percentile is between 89 and 95 per cent and the width is between 45 and 115 Å the index varies by less than 1 per cent. These results are independent of our choice of test spectra. We therefore define the

![Figure 4](https://academic.oup.com/mnras/article-abstract/445/3/2575/1044976)

**Figure 4.** The value of the index (colour) as a function of both the effective temperature and the metallicity for all dwarfs in the synthetic grid at [$\alpha$/Fe] = 0.00 dex.

![Figure 5](https://academic.oup.com/mnras/article-abstract/445/3/2575/1044976)

**Figure 5.** The mean and spread in the index over the full range in stellar parameters as a function of SNR.
pseudo-continuum at any wavelength as the 90th percentile flux in 100 Å window around that wavelength.

4 VALIDATION

To achieve a good validation we needed several validation sets. The various data sets were employed to allow us to study the performance of our method in different regimes of stellar parameter space, across large data sets and with low-resolution SDSS data. We describe the validation sets and the reason they were chosen separately below.

4.1 Comparison with the ELODIE LIBRARY

The ELODIE library is a publicly available library of about 2000 spectra of 1388 stars observed with the ELODIE spectrograph on the Observatoire de Haute-Provence 1.93-m telescope. ELODIE 3.1 (Prugniel & Soubiran 2001; Mould et al. 2004) contains high-resolution spectra ($R = 42000$), high SNR (>100) observations of stars with 3700 K < $T_{\text{eff}}$ < 13 600 K, 0.03 dex < log g < 5.86 dex and −2.8 dex < [Fe/H] < +0.17 dex. Lee et al. (2011) used a sample of 425 ELODIE spectra that have well-known stellar parameters and their [α/Fe] from the literature. It is important to note that because the [α/Fe] values are taken from the literature there may be some scatter that exist in this validation set (Lee et al. 2011).

We processed the ELODIE spectra in the same way as the rest of the spectra: degrading the resolution to $R = 2000$ using a Gaussian kernel, resampling the spectra to 0.5 Å pixels, applying a BMC normalization. We calculated the index for all stars that fall in our parameter range and studied the index as a function of [α/Fe] (Fig. 7). The colour scheme is the same as Fig. 3.

We see a tight correlation between the index and the α-abundance over the full range of stellar parameters indicating that it is possible to estimate the [α/Fe] directly from the index. This result shows us that we can rank stars from low to high [α/Fe] using the index only. A star with an index below approximately 7.2 would be considered α-poor while a star with an index about 7.4, α-rich. These index values are consistent with the synthetic spectra in Section 3.4. Although a large number of stars exist in the ELODIE library, they are mostly metal-rich compared to the halo stars we are interested in. As such, this validation set is not completely suitable for our purposes. Thus, we must explore other validation sets.

4.2 Comparison with Nissen & Schuster data

N10 measured the Mg, Ti, Si, Ca and Fe-peak elemental abundance abundances of 78 halo stars and 16 disc stars using spectra from the Very Large Telescope’s (VLT) Ultraviolet and Visual Echelle Spectrograph (UVES) spectra and the Fibre fed Echelle Spectrograph (FIES) on the Nordic Optical Telescope (NOT). These spectra were kindly provided to us by P. Nissen. The wavelength coverage of VLT/UVES sample extends only to ∼4700 Å. Since the index includes a Ti i feature at ∼4500 Å, the VLT/UVES data from N10 could not be used for validation. The NOT/FIES spectra have a wavelength coverage that allows us to use this Ti i feature. In total, there are 47 stars from the N10 data set that meet our criteria in wavelength coverage and stellar parameters space. N10 represents the parameters space we are most interested in, and is also internally consistent. This data set is rather small and thus this set alone does not allow detailed analysis of the performance of the index but complements the ELODIE validation set. Fig. 8 shows the performance of the index as a function of [α/Fe] for the N10 data. The colour coding represents the stellar parameters as in Fig. 3.

We found again a tight correlation between the index and the α-abundance over the full range of stellar parameters. Fig. 8 makes it clear that we can use the index alone to categorize the ‘high-α’ and ‘low-α’ stars from N10. Stars with an index below 7.2 would be considered α-poor while a star with an index of 7.4 is α-rich. These values are the same as for the synthetic grid and the ELODIE validation sets. The tightness of the linear correlation between the index and the [α/Fe] is further proof that we may be able to estimate the α-abundance directly from the index.

4.3 Comparison with SDSS calibration targets

Finally, we used a sample of SDSS spectra with accurate stellar parameters and [α/Fe] determined from high-resolution spectra. This sample was used in Lee et al. (2011) and was originally constructed to validate the SSPP pipeline (for more details consult Allende Prieto et al. 2008; Smolinski et al. 2011). We used a subset, totalling 73 stars, from the SDSS calibration targets discussed in Lee et al. (2011) that exist inside our stellar parameter range. We studied the index as a function of [α/Fe] and the stellar parameters in Fig. 9. The colour scheme and axes are the same as Fig. 3. We see again that it is possible to separate stars based on their α-abundances using the index alone as long as the star exists inside our parameter range. The index value needed to separate stars into α-rich and α-poor populations is consistent with the tests on the synthetic grid as well as the ELODIE and N10 validation sets.

4.4 Converting the Index to an Estimate of [α/Fe]

We have shown, for all of our tests with synthetic and real spectra, that there is a linear relationship between our index and [α/Fe]. We used a linear regression model to approximate the index as a function of [α/Fe], to estimate the [α/Fe] directly from the index. The ELODIE validation set was used as a calibration set because it is the largest one. A linear regression model for the ELODIE data set yields

$$\text{Index} = [\alpha/\text{Fe}] \times 4.32 + 6.28. \quad (2)$$

Figure 6. Stability of index as a function of BMC parameters. ‘X’ is the location suggested by Rogers et al. (2010). The shaded region represents the region where the index varies by an order of less than 1 per cent.
The regression model achieved a correlation coefficient of 0.86 and is statistically significant with relatively low scatter (5 per cent in index space). The equation to convert our index into a direct estimate of $\alpha$/Fe was tested on the remaining validation sets (N10, and SDSS high-resolution targets). We also added an additional validation set using globular clusters and open clusters to compare the estimated $\alpha$/Fe with the high-resolution average measured in the literature.

For the case of N10, we used the ELODIE data set as a calibration set to convert our index in the $\alpha$/Fe through equation (2), we found the $\alpha$/Fe and estimated $\alpha$/Fe are in good agreement with a mean offset of 0.01 dex and an external error of ±0.05 dex (Fig. 10).
We also found no significant offsets or correlations between the residuals in [$\alpha$/Fe] and the stellar parameters (Fig. 11). Finally, we repeated the analysis for the SDSS calibration targets and found a very good agreement between our estimated $\alpha$-abundances and the high-resolution estimates with an offset in [$\alpha$/Fe] = 0.00 dex and an external error of $\pm$0.10 dex.

Figure 9. The index value as a function of [$\alpha$/Fe] broken down by $T_{\text{eff}}$ (top panel), [Fe/H] (middle panel), $\log g$ (bottom panel) for 73 SDSS stars with parameters obtained by high-resolution spectra that exist in our parameter range. The error bar represents the typical errors on the [$\alpha$/Fe] from the SDSS high-resolution validation set.

Figure 10. The estimated [$\alpha$/Fe] as a function of the high-resolution [$\alpha$/Fe] measurement from N10. The blue line represents a 1:1 relation. The error bar represents the typical errors on the [$\alpha$/Fe] from the N10 validation set.

Figure 11. The residual between the estimated and true $\alpha$/Fe] broken down by $T_{\text{eff}}$ (top panel), $\log g$ (middle panel), [Fe/H] (bottom panel) for the NOT/FIES subset of the N10 data.
Table 2. Globular/open cluster data.

| Cluster | ([Fe/H])_{lit} (dex) | ([α/Fe])_{lit} (dex) | ([α/Fe])_{est} (dex) | N_{mem} |
|---------|---------------------|---------------------|---------------------|--------|
| M13     | -1.58 ± 0.04        | 0.20 ± 0.04         | 0.19 ± 0.05         | 57     |
| M71     | -0.80 ± 0.02        | 0.23 ± 0.10         | 0.13 ± 0.05         | 5      |
| NGC 2420| -0.05 ± 0.02        | 0.03 ± 0.09         | 0.08 ± 0.08         | 72     |

Note. The data for each of the clusters were compiled from the following sources: M13: Sneden et al. (2004), Cohen & Meléndez (2005); M71: Boesgaard et al. (2005); NGC2420: Pancino et al. (2010). N_{mem} is the total number of stars in the SDSS that were used to compute the average abundances.

4.5 Comparison with SDSS clusters

We used a sample of two globular clusters and one open cluster observed with the SDSS to validate our estimated [α/Fe] values. Clusters are a good test bed to further validate our method, as each cluster should have an average [α/Fe] with a relatively small spread. While clusters have small star-to-star variations in their abundances of light elements (C, N, Al, Na) we expect these to not strongly affect the α-elements and thus the [α/Fe] measurement. In order to achieve a large enough sample of stars in the clusters, we had to decrease the log g constraint to log g > 3.0 dex. Further, we require the SNR to be larger than 40 (see Section 3.6) to only select out stars for which we can obtain suitable values of [α/Fe].

The literature values for [α/Fe] were compiled by averaging the mean cluster abundances of the individual α-elements for each cluster. The standard deviation was also computed and is displayed in Table 2. We estimated the [α/Fe] for likely cluster members from M71 and M15, and the open cluster NGC 2420 using our SIM. The other clusters of Lee et al. (2011), M13, M67, NGC 6791, are outside our metallicity range. For the three clusters that are in our parameter range, we estimated the mean and standard deviation in [α/Fe]. Both the literature and estimated values for the mean [α/Fe] can be found in Table 2. Our estimated mean [α/Fe] for each cluster is in very good agreement with the literature values providing further proof that our method can produce meaningful values for [α/Fe].

4.6 Computation of internal and external error on the index

The total internal error on the index at any SNR can be estimated by completing noise injection experiments on both the SDSS data and synthetic spectra. We do the later in Section 3.6 and find the internal error can be as high as 0.60 (≈10 per cent) at SNR = 40. Propagating this to estimate the uncertainty in [α/Fe] via equation (2) yields $\sigma_{\text{internal}} = \pm 0.13$ dex. The external uncertainty in the [α/Fe] is estimated by using the validation sets is shown to be as high as $\sigma_{\text{external}} = \pm 0.1$ dex. Thus, a conservative estimate of the total uncertainty in [α/Fe] is the internal and external errors added in quadrature and is $\sigma_{[\alpha/Fe]} = \pm 0.16$ dex at a SNR = 40. The external and internal errors are comparable but the external error can be greatly decreased by carefully selecting the range of stellar parameters. For example, when we constrained the temperature between 6000 and 6500 K and the metallicity between −0.8 and −2.0 dex, i.e. the expected turnoff region in our sample, we found that the index varies of the order of 0.20. This implies we can estimate [α/Fe] with an internal error of ±0.05 dex by constraining the stellar parameter to focus on the turnoff region only.

Our method improves on the current methods of Lee et al. (2011) and Franchini et al. (2011) by decreasing the dependence on models and uncertainties in the stellar parameters. We have shown that with our empirical SIM we can estimate [α/Fe] to within ~±0.1 dex when accounting for the full parameter space described in Section 3.4, which has a temperature span of 2000 K, log g span of 1.5 dex and a metallicity coverage of 2.0 dex. In context, the spectral matching technique of Lee et al. (2011) has an internal error of ~±0.1 dex in [α/Fe] for a temperature uncertainty of only ±300 K.

5 RESULTS

5.1 Distribution of α-elements in the inner halo

The effective temperature and metallicity for each of our targets are sourced from the adopted values from the SDSS SSP. We computed the [α/Fe] for our targets using the SIM. The estimated [α/Fe] distribution and metallicity-[α/Fe] diagram are shown in Figs 12(a) and (b), respectively. The peak of the [α/Fe] in Fig. 12(a) is +0.30 dex which confirms that halo stars are, on average, α-enhanced with a large sample from SDSS (Helmi 2008; Feltzing & Chiba 2013, and references therein). The dash–dotted line in Fig. 12(a) shows the Gaussian fit to the [α/Fe] distribution function. It is interesting to point out that the fit over predicts the amount of low-α stars near [α/Fe] = 0.20 dex. We plan to explore this and the relative ratio of these two populations as a function of Galactic parameters further in the next work of this series. Constraining the sample to just the turnoff region, i.e. 6000 < T_{eff} < 6500 K, log g > 3.5 dex, −0.8 < [Fe/H] < −2.0 dex, produced a standard deviation of the distribution in [α/Fe] of 0.15 dex. This dispersion is larger than the estimated uncertainty in the constrained parameter space indicating we can resolve the α-rich and α-poor populations.

Motivated by the conversion described in Section 4.4 and N10, α-poor stars are defined as any stars which have an index less than 6.93, corresponding to an [α/Fe] < +0.15 dex and α-rich stars as any stars which have an index larger than 7.79, corresponding to an [α/Fe] > +0.35 dex. These criteria are designed to account for the boundary of α-rich and α-poor stars at ~[α/Fe] = +0.25 dex (N10) and the uncertainty in the estimated [α/Fe]. The criteria are used to select, in a statistical way, stars which are have a high probability of being either α-rich or α-poor based on the 1σ uncertainties in the estimated [α/Fe].

5.2 Turnoff detection and its uncertainties

The turnoff temperature can be used to determine the age of the youngest stellar population (for further discussion on this method consult Soderblom 2010). We used a Sobel–Kernel edge detector algorithm to determine the temperature of the main-sequence turnoff (for more details consult Jofr & Weiss 2011). The method assumes the temperature distribution of a stellar population will display a very sharp decline (i.e. edge) near the turnoff as the more massive, hotter stars in the population have evolved off the main sequence. The edge, or turnoff temperature in our case, was determined by constructing a temperature distribution function for a range of metallicity bins and computing maximum of its derivative. This algorithm is sensitive to the sampling and in order to determine the uncertainty in the turnoff temperature we used a bootstrap method (Tabur, Kiss & Bedding 2009; Jofr & Weiss 2011). In the
The results of this simulation confirm that if the bin size is smaller than $\sim 2\sigma_{[\text{Fe}/\text{H}]}$ then an individual star in our sample can jump between different metallicity bins and thus change the detected turnoff temperature significantly (Jofrè & Weiss 2011). On the other hand, excessively large bin sizes can lead to a less precise turnoff temperature because the turnoff temperature is metallicity-dependent. Typical uncertainties in [Fe/H] at these higher SNR are of the order of 0.15–0.2 dex. Our simulations showed that the uncertainty in the turnoff temperature due to the metallicities becomes small relative to the bootstrapping error as the metallicity bin size becomes larger than $\sim 0.3$ dex. Therefore, the uncertainty in the turnoff temperature is defined as the $3\sigma$ error as a result of the bootstrapping with a bin size in metallicity of 0.3 dex.

5.3 Metallicity–temperature diagram

In the left-hand panel of Fig. 13, we plot the turnoff temperature as a function of metallicity of our SDSS sample overplotted with the $\alpha$-rich and $\alpha$-poor Dartmouth (Dotter et al. 2008) isochrones. The right-hand panel of Fig. 13 shows the turnoff temperature as a function of metallicity for our sample overplotted with the Yonsei–Yale ($Y^2$; Demarque et al. 2004) isochrones. The first thing to note is that the turnoff temperature as a function of metallicity does not follow along an isochrone of a constant age indicating there is a correlation between age and metallicity. These plots also show that the turnoff temperature of the $\alpha$-rich and $\alpha$-poor stars in our sample are, within the errors, equal at low metallicity. As the metallicity increases, upwards of $\sim 1.4$ dex the $\alpha$-rich stars have a significantly lower turnoff temperature indicating they are older than the $\alpha$-poor stars. All of our stars are older, 8 Gyr with both isochrones. The Dartmouth isochrones produce larger ages compared to the $Y^2$ isochrones and in some cases have ages larger than the accepted age of the Universe (13.8 Gyr; Planck Collaboration I et al. 2013). This is likely a result of the prescription of atomic diffusion (Jofrè & Weiss 2011).

Qualitatively, the turnoff temperature of the $\alpha$-rich component as function of metallicity is not significantly different with the results of Jofrè & Weiss (2011). This is expected as the bulk of the stars in the sample are $\alpha$-rich. Another interesting finding we observe from the overplotted isochrones is that the Dartmouth isochrones have larger differences in the turnoff temperature between $\alpha$-rich and $\alpha$-poor stars at the same age. This indicates that the two isochrones have different responses to $\alpha$-enhancement. The different isochrone sets also use different helium mixtures, which can also affect the isochrones.

5.4 Isochrone analysis: ages and their errors

With the turnoff temperatures, metallicities and a set of isochrones in hand, we can determine the age of our $\alpha$-rich and $\alpha$-poor populations. For this, we used the $Y^2$ isochrones (Demarque et al. 2004) which were interpolated to determine the theoretical turnoff temperature as a function of age, metallicity and $[\alpha/\text{Fe}]$. The $Y^2$ isochrones include atomic diffusion (Demarque et al. 2004) while the Dartmouth isochrones only include moderate atomic diffusion (Dotter et al. 2008). As a result, the Dartmouth isochrones produce larger ages compared to the $Y^2$ isochrones. To avoid dependencies in the isochrones used, we have considered the relative ages between the $\alpha$-rich and $\alpha$-poor populations.

The effect of increasing $[\alpha/\text{Fe}]$ in an isochrone model, while keeping [Fe/H] fixed, is to effectively increase the total metal abundance. Dartmouth and $Y^2$ isochrones show that increasing the $\alpha$-abundance...
Ages of two $\alpha$-populations in the halo

6 DISCUSSION

6.1 Age–metallicity relation

We have, for the first time, measured the age–metallicity relation of an $\alpha$-rich and $\alpha$-poor population in the Galactic halo at metallicities between $-0.80$ and $-2.0$ dex (Fig. 14). We note that: (1) at high metallicities ([Fe/H] $>-1.4$ dex), the $\alpha$-rich population is older than the $\alpha$-poor population, and (2) at low metallicities ([Fe/H] $<-1.4$ dex), the two populations are coeval within the errors. We have found that the shape of the age–metallicity relation differs for both $\alpha$-populations.

The $\alpha$-poor stars tend to have a steeper correlation between age and metallicity compared to the $\alpha$-rich stars. This is consistent with the idea that the $\alpha$-poor stars were formed in areas with slower chemical evolution compared to their $\alpha$-rich counterparts. The very shallow slope on the age–metallicity correlation for $\alpha$-rich stars seems to infer they were formed in a quick event with high star formation rates which were able to produce a broad range in metallicities in a short time (less than $\sim10^9$ yr) to keep the $\alpha$-abundance enhanced.

By testing multiple isochrone sets, we have found that our age sequence at low metallicities, namely that the $\alpha$-rich population is coeval with the $\alpha$-poor population, within the errors, is robust. Our age difference between the $\alpha$-rich and $\alpha$-poor populations in the most metal-rich bin is affected by as much as 1 Gyr. However, this difference does not effect the interpretation of our results.

It is also worth commenting that the age–metallicity relation that we observe in Fig. 14, is similar to the ‘Y’-like bifurcation seen in the globular clusters of Marín-Franch et al. (2009) and Leaman, VandenBerg & Mendel (2013). Marín-Franch et al. (2009) found two well-defined tracks in the age–metallicity diagram for globular...
clusters: a ‘young’ track which can be associated with dwarf galaxies and an ‘older’ track with a small age dispersion which can be associated with an in situ population that may have formed from a protogalactic gas cloud. We found a similar bifurcation in our halo field star sample which supports a scenario of common origin for the GC and halo field stars (e.g. Martell & Grebel 2010).

6.2 Implications for the formation of the Galactic halo

One unambiguous result from our study is that all of the stars in our sample, regardless of which isochrone set is used, are older than \( \sim 8 \) Gyr. This confirms the notion that the Galactic halo was formed and assembled very early on as suggested by both theory (Robertson et al. 2005; Font et al. 2006, 2011; Zolotov et al. 2009, 2010) and observations (e.g. Schuster et al. 2012).

The theoretical work of Robertson et al. (2005) and Font et al. (2006) lays out what Schuster et al. (2012) refer to as a ‘dual-accretion’ scenario. In this scenario, the halo is built from the bottom-up via hierarchical merging. The dual-accretion models produce a Galactic halo by merging smaller subunits of stars together rapidly at early times fully destroying the subunits in the process. The protohalo systems form massive stars rapidly and are enriched primarily by Type II supernova producing an \( \alpha \)-rich population. Since the subunits are broken up quickly, there is little iron enrichment via Type Ia supernova leaving behind a set of low metallicity, \( \alpha \)-rich stars. These would correspond to the \( \alpha \)-rich stars in our sample. Longer-lived dwarf galaxies, on the other hand, are polluted via both types of supernova leading to a population of stars that have a wide range in metallicities and are \( \alpha \)-poor. These stars are accreted on to the Milky Way at later times and could represent the \( \alpha \)-poor stars in our sample.

Interestingly, the age–metallicity diagram from the theoretical work of Font et al. (2006) appears to have two tracks, one that has a steeper slope in age–metallicity space than the other which is the case in our sample (Fig. 14). This bifurcation in the age–metallicity relation can be interpreted as evidence that the Galactic halo was built in two phases. One quick phase that produced the \( \alpha \)-rich stars over broad metallicity ranges, leading to a shallow age–metallicity relation. Followed by a second phase of accretion of stars that evolved in environments of slower chemical evolution, leading to a more pronounced age–metallicity relation.

The more recent hydrodynamic simulations by Font et al. (2011) have predicted that the accreted stars are, on average, 3–4 Gyr older than the in situ population. If we make the assumption that the in situ population are tracked by our \( \alpha \)-rich stars (an assumption often made in the literature; e.g. Schuster et al. 2012) than our relative age sequence would indicate that the accreted stars are \( \sim 1 \) Gyr younger than the in situ population. This result does not agree with the results of Font et al. (2011) and is only about half 2–3 Gyr age difference observed by Schuster et al. (2012). However, this discrepancy could be remedied if we do not make the assumption that \( \alpha \)-rich stars are the in situ population and the \( \alpha \)-poor stars are the accreted population or if the feedback in their simulations were less efficient in low-mass haloes (Font et al. 2011). The unavoidable (larger) errors on \([\alpha/Fe]\) with low-resolution spectra can lead to a slight mixing of the two populations which could also explain the diminished age difference compared to the results of Schuster et al. (2012).

Contrary to Robertson et al. (2005) and Font et al. (2006, 2011), the models of Zolotov et al. (2009, 2010) predicted that the in situ population within the Galactic halo is likely formed within inner 4 kpc of the Galaxy’s centre and is formed via cold, infalling gas. The in-falling gas creates a primeval bulge or disc. After which those stars in the primeval disc are kinematically heated into the halo via early accretion and merging events which at the same time populate the accreted component of the halo via tidal stripping (Purcell, Bullock & Kazantzidis 2010). These in situ stars can be interpreted as our \( \alpha \)-rich stars (just as in Schuster et al. 2012). Stars formed in the dwarf galaxies that later merged with the primeval disc/bulge can be attributed to our \( \alpha \)-poor stars. Zolotov et al. (2010) found, in all but one of their model haloes, the in situ population was formed early on and the accreted population was formed, on average, at later times and then subsequently accreted on to the halo. However, it is not clear whether this is true at all metallicities or just higher metallicities. Our relative age sequence at high metallicities, namely \( \alpha \)-rich stars being older than \( \alpha \)-poor stars on average, favours the models of Zolotov et al. (2009, 2010) and at high metallicities confirms the observational results of Schuster et al. (2012) with a larger, statistical sample. Further the bifurcation we observed should still be present in these models. The results of this study can be used to help construct more realistic models and better constrain the basic physics, in particular feedback mechanisms, which can replicate the observations.

7 SUMMARY

Motivated by the results of both Schuster et al. (2012) and Jofré & Weiss (2011), we addressed the intriguing problem of the relative ages of the \( \alpha \)-rich and \( \alpha \)-poor stars in the Galactic halo with a large sample of stars from SDSS. To that end, we developed a new spectral index-based method to estimate the \([\alpha/Fe]\) abundances using low-resolution (\( R \sim 2000 \)), moderate SNR (\( >40 \)) SDSS spectra. We studied the ages of a sample of main-sequence turnoff halo field stars selected using a cuts in colour and Galactic latitude. With our method (described in Section 3), we could split the halo field star population statistically into an \( \alpha \)-rich and \( \alpha \)-poor subsample. We used a Sobel–Kernel edge detection method to determine the turnoff temperature, and thus ages of the two stellar subsamples. Our results can be summarized in the following points.
(i) A spectral index-based method (see Section 3 for more details) was constructed to rank stars based on $\alpha$/Fe. We have estimated the uncertainties of the $\alpha$/Fe to be of the order of $\pm 0.15$ dex. This method can be used on a range of stellar parameters and is a semiempirical method to estimate $\alpha$/Fe that automatically accounts for the stellar parameters. The uncertainty in $\alpha$/Fe from our method is comparable to the other methods (e.g. Lee et al. 2011). Using this method, we found that the halo is comprised of an $\alpha$-rich and $\alpha$-poor population which may peak near $\alpha$/Fe $\sim$ +0.40 and +0.20 dex, respectively (see Fig. 12). It may be possible to extend this basic method to find stars with spectral peculiarities (e.g. s-process enriched star, etc.) which is something we plan to explore in future. Finally, this method can be expanded to other elemental species and spectral-types.

(ii) Given the large absolute age, based on the $Y^2$ isochrones which included $\alpha$-enrichment and atomic diffusion, the Galactic halo must have been formed very early on. It is also likely that the Galaxy is a relatively quiet place not having undergone a major merging event in the last 8 Gyr. This is consistent with other observations and theoretical studies (e.g. Robertson et al. 2005; Font et al. 2006; Zolotov et al. 2009, 2010; Schuster et al. 2012).

(iii) We have made a first measure of the age–metallicity relation of halo field stars separated by $\alpha$-abundances at low metallicities. There appears to be a difference, of the order of 1 Gyr, in the ages of the $\alpha$-rich and $\alpha$-poor subsamples (see Section 5). Using the $Y^2$ isochrones, we found the $\alpha$-rich subsample is older than the $\alpha$-poor subsample in the high-metallicity case ([Fe/H] $\gtrsim$ -1.4 dex) which confirms the observational results of Schuster et al. (2012) and the theoretical results of Zolotov et al. (2009, 2010). However, we extended the results of Schuster et al. (2012) to lower metallicities and found there is a break around [Fe/H] $\sim$ -1.4 dex, and the $\alpha$-rich subsample becomes coeval with the $\alpha$-poor subsample. Interestingly, this bifurcation in the age–metallicity diagram is also seen in globular clusters. This hints that the $\alpha$-rich population, with a shallower age–metallicity relation, was formed in a rapid event with high star formation (e.g. collapse of a protogalactic gas cloud), while the $\alpha$-poor stars were formed in an environment with a slower chemical evolution time-scale. Our results also confirm the idea that the $\alpha$-rich subsample may be a population formed in situ while the $\alpha$-poor subsample may have formed in satellite galaxies and accreted into the Milky Way halo.

The relative age sequence we found supports the models of Zolotov et al. (2009, 2010). These models point out that we can use the fraction of in situ and accreted stars to disentangle the importance of accretion events in the assembly of the Galactic halo. With this new spectral index-based method to estimate $\alpha$/Fe from low-resolution spectra, we are in a good position to study the relative fraction of $\alpha$-rich and $\alpha$-poor stars in the Galactic halo.

With upcoming large-scale surveys such as Gaia (Perryman et al. 2001), Gaia-ESO (Gilmore et al. 2012), 4MOST (de Jong et al. 2012), and GALAH (Anguiano et al. 2014), and it will be possible to resolve large sample of stars in position, velocity, distance and chemical phase-spaces. The next generation of data will undoubtedly open new windows into studying the ages of individual halo field stars as a function of chemical abundances providing detailed insights of the early stages of our home Galaxy.

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REFERENCES

Ahn C. P. et al., 2012, ApJS, 203, 21
Allende Prieto C. et al., 2008, AJ, 136, 2070
Alvarez R., Plez B., 1998, A&A, 330, 1109
Anguiano B. et al., 2014, in Feltzing S., Zhao G., Walton N. A., Whitelock P., eds, Proc. IAU Symp. 298, Setting the Scene for Gaia and LAMOST. Cambridge Univ. Press, Cambridge, p. 322
Asplund M., Grevesse N., Sauval A. J., 2005, in Barnes T. G., III Bash F. N., eds, ASP Conf. Ser. Vol. 336, Cosmic Abundances as Records of Stellar Evolution and Nucleosynthesis. Astron. Soc. Pac., San Francisco, p. 25
Barklem P. S., O’Mara B. J., 1998, MNRAS, 300, 863
Belokurov V. et al., 2006, ApJ, 642, L137
Boesgaard A. M., King J. R., Cody A. M., Stephens A., Deliyannis C. P., 2005, ApJ, 629, 832
Carollo D. et al., 2007, Nature, 450, 1020
Carollo D. et al., 2010, ApJ, 712, 692
Cohen J. G., Meléndez J., 2005, AJ, 129, 303
de Jong R. S. et al., 2012, in McLean I. S., Ramsay S. K., Takami H., Proc. SPIE Conf. Ser. Vol. 8446, Ground-based and Airborne Instrumentation for Astronomy IV. SPIE, Bellingham, p. 84460T
Demarque P., Woo J.-H., Kim Y.-C., Yi S. K., 2004, ApJS, 155, 667
Dotter A., Chaboyer B., Jevremović D., Kostron D., Baron E., Ferguson J. W., 2008, ApJS, 178, 89
Eggen O. J., Lynden-Bell D., Sandage A. R., 1962, ApJ, 136, 748
Feltzing S., Chiba M., 2013, New Astron. Rev., 57, 80
Font A. S., Johnston K. V., Bullock J. S., Robertson B. E., 2006, ApJ, 638, 585
Font A. S., McCarthy I. G., Crain R. A., Theuns T., Schaye J., Wiersma R. P. C., Dalla Vecchia C., 2011, MNRAS, 416, 2802
Francini M., Morossi C., Di Marcantonio P., Malagnini M. L., Chavez M., 2010, ApJ, 719, 240
Francini M., Morossi C., Di Marcantonio P., Malagnini M. L., Chavez M., 2011, ApJ, 730, 117
Gilmore G. et al., 2012, The Messenger, 147, 25
Gustafsson B., Edvardsson B., Eriksson K., Jørgensen U. G., Nordlund Å., Plez B., 2008, A&A, 486, 951
Helmi A., 2008, A&AR, 15, 145
Hill V. et al., 2002, A&A, 387, 560
Ibata R. A., Gilmore G., Irwin M. J., 1994, Nature, 370, 194
Ivezić Z., Beers T. C., Jurić M., 2012, ARA&A, 50, 251
Jofrè P., Weiss A., 2011, A&A, 533, A59
Kalirai J. S., 2012, Nature, 486, 90
Kim Y.-C., Demarque P., Yi S. K., Alexander D. R., 2002, ApJS, 143, 499
Kordopatis G. et al., 2013, MNRAS, 436, 3231
Kupka F., Ryabchikova T. A., 1999, Publ. Obs. Astron. Beograd, 65, 223
Leaman R., VandenBerg D. A., Mendel J. T., 2013, MNRAS, 436, 122
Lee Y. S. et al., 2008a, AJ, 136, 2022
Lee Y. S. et al., 2008b, AJ, 136, 2050
Lee Y. S. et al., 2011, AJ, 141, 90
McWilliam A., 1997, ARA&A, 35, 503
Marín-Franch A. et al., 2009, ApJ, 694, 1498
Martell S. L., Grebel E. K., 2010, A&A, 519, A14
Masseron T., 2006, PhD thesis, Observatoire de Paris
Masseron T. et al., 2014, in press
Mouldtaka J., Ilovaisky S. A., 2014, PASP, 126, 693
Nissen P. E., Schuster W. J., 2010, A&A, 511, L10 (N10)
Nissen P. E., Schuster W. J., 2011, A&A, 530, A15
Nissen P. E., Schuster W. J., 2012, A&A, 543, A28
Pancino E., Carrera R., Rossetti E., Gallart C., 2010, A&A, 511, A56
Perryman M. A. C. et al., 2001, A&A, 369, 339
Planck Collaboration I 2013, preprint (arXiv:1303.5062)
Plez B., 2012, Astrophysics Source Code Library, record ascl:1205.004
Prugniel P., Soubiran C., 2001, A&A, 369, 1048
Purcell C. W., Bullock J. S., Kazantzidis S., 2010, MNRAS, 404, 1711
Ramírez I., Meléndez J., Chanamé J., 2012, ApJ, 757, 164
Robertson B., Bullock J. S., Font A. S., Johnston K. V., Hernquist L., 2005, 
ApJ, 632, 872
Rogers B., Ferreras I., Peletier R., Silk J., 2010, MNRAS, 402, 447
Schönrich R., Asplund M., Casagrande L., 2011, MNRAS, 415, 3807
Schuster W. J., Moitinho A., Márquez A., Parrao L., Covarrubias E., 2006, 
A&A, 445, 939
Schuster W. J., Moreno E., Nissen P. E., Pichardo B., 2012, A&A, 538, A21
Searle L., Zinn R., 1978, ApJ, 225, 357
Sheffield A. A. et al., 2012, ApJ, 761, 161
Smolinski J. P. et al., 2011, AJ, 141, 89
Sneden C., Kraft R. P., Guhathakurta P., Peterson R. C., Fullbright J. P., 2004, 
AJ, 127, 2162
Soderblom D. R., 2010, ARA&A, 48, 581
Tabur V., Kiss L. L., Bedding T. R., 2009, ApJ, 703, L72
Thomas D., Maraston C., Bender R., 2003, MNRAS, 339, 897
Unavane M., Wyse R. F. G., Gilmore G., 1996, MNRAS, 278, 727
Yanny B. et al., 2009, AJ, 137, 4377
York D. G. et al., 2000, AJ, 120, 1579
Zolotov A., Willman B., Brooks A. M., Governato F., Brook C. B., Hogg 
D. W., Quinn T., Stinson G., 2009, ApJ, 702, 1058
Zolotov A., Willman B., Brooks A. M., Governato F., Hogg D. W., Shen S., 
Wadsley J., 2010, ApJ, 721, 738

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