Clinical characteristics of patients with \textit{ROS1} gene rearrangement in non-small cell lung cancer: a meta-analysis

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\textbf{Background:} \textit{ROS1} gene rearrangement has been reported in several types of cancers, including non-small cell lung cancer (NSCLC). It is reported that tyrosine kinase inhibitors are effective in the treatment of \textit{ROS1}-rearranged NSCLC. Therefore, the identification of \textit{ROS1} rearrangement can be used as potential therapeutic target in lung cancer. Epidemiological data indicates that \textit{ROS1} gene rearrangement occurs in approximately 1–2% of NSCLC patients. The small sample sizes of the existing associated studies only represent the characteristics of patients in specific regions or countries, and there is still no latest statistical analysis on \textit{ROS1} gene rearrangement anywhere in the world.

\textbf{Methods:} We conducted a systematic search of the PubMed, Embase, Cochrane Central Register of Controlled Trials (CENTRAL), CBM, CNKI, Wanfang, and VIP databases to identify studies on \textit{ROS1} gene rearrangement in NSCLC patients from January 1, 2015 to October 27, 2019. We conducted a meta-analysis to investigate the relationship between \textit{ROS1} gene rearrangement and clinical characteristics of NSCLC patients. The four clinical features are as follows: gender, smoking status, pathological type, and lung cancer stage.

\textbf{Results:} Thirty-nine studies constituting of 25,055 NSCLC patients were eligible for inclusion in this meta-analysis. A prominently higher rate of \textit{ROS1} gene rearrangement was observed in female NSCLC patients (OR =1.94, 95% CI: 1.62–2.32%, \textit{P}<0.05), patients with no smoking history (OR =2.82, 95% CI: 2.24–3.55%, \textit{P}<0.05), patients with adenocarcinoma (OR =1.55, 95% CI: 1.14–2.11%, \textit{P}<0.05), and patients with stage III–IV disease (OR =1.50, 95% CI: 1.15–1.94%, \textit{P}<0.05). Our meta-analysis also showed that the prevalence of \textit{ROS1} rearrangement in adenocarcinoma was 2.49% (95% CI: 1.92–3.11%), while it was lower in non-adenocarcinoma patients (1.37%).

\textbf{Conclusions:} \textit{ROS1} gene rearrangement was more predominant in female patients, patients without smoking history, patients with adenocarcinoma and patients with advanced-stage disease (stages III to IV).

\textbf{Keywords:} \textit{ROS1}; clinicopathologic features; non-small cell lung cancer (NSCLC); meta-analysis

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\section*{Introduction}

Non-small cell lung cancer (NSCLC) accounts for 85% of all lung cancer cases (1). The majority of NSCLC patients are locally advanced or metastatic at the time of diagnosis which leads to a dire overall prognosis (2). In the past decade, the clinical application of targeted therapy and immuno therapy has broken the traditional single-treatment model of lung cancer, and the 5-year survival rate has risen to 15.5% (3).
In the age of precision medicine, molecular detection has become extremely important way of identifying the type of gene mutations in lung cancer. In normal lung tissue, **ROS1**, a receptor tyrosine kinase (RTK), is usually expressed at low levels (4). The **ROS1** gene was first discovered in 1986 in ornithon sarcoma virus (UR2), which has a unique carcinogenic gene sequence (5). The **ROS1** gene is located in the q12 region of chromosome 6, encoding 2,347 amino acids with a molecular weight of 259 kDa (6,7). In 1987, **FIG-ROS1** fusion gene was first discovered in pleomorphic glioblastoma cell line U118MG (8). In fact, the expression of **ROS1** fusion protein is not well understood, however, the PIK/AKT, MAPK/ERK, and JAK/STAT3 signaling pathways are known to be related (13). **ROS1** bears some resemblance to the **ALK** (14); therefore, **ALK** inhibitor crizotinib is regarded as the first-line therapy for **ROS1**- rearranged patients (15). These promising results emphasize the necessity of investigating **ROS1** gene rearrangement in patients with NSCLC. Because the probability of **ROS1** gene rearrangement is low, and most of the sample sizes are small to medium, it is not enough to fully represent the current state of **ROS1** gene rearrangement mutation.

In addition, we present the following article in accordance with the PRISMA reporting checklist (http://dx.doi.org/10.21037/tcr-20-1813).

**Methods**

**Literature search**

We performed our meta-analysis according to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines (16). Articles from between January 1, 2015 and October 27, 2019, were systematically retrieved from the PubMed, Embase, Cochrane Central Register of Controlled Trials (CENTRAL), CBM, CNKI, Wanfang, and VIP databases. Search terms were retrieved as title/abstract by MESH subject word and customary terminology, the final retrieval strategy is formulated after several pre-searches, and the search records are output to Endnote. The search term is Carcinoma, Non-Small-Cell Lung/Carcinomas, Non-Small-Cell Lung/Lung Carcinoma, Non-Small-Cell/Lung Carcinomas, Non-Small-Cell/Non-Small-Cell Lung Carcinomas/Nonsmall Cell Lung Cancer/Non-Small-Cell Lung Carcinoma/Non Small Cell Lung Carcinoma/Carcinoma, Non-Small Cell Lung/Non-Small Cell Lung Cancer/non-small cell lung cancer/NSCLC/ lung neoplasm/lung cancer/lung carcinoma/pulmonary neoplasm/pulmonary cancer/pulmonary carcinoma; **ROS1**/ **ROS-1/c-ros protein/Ros1 proto-oncogene protein/Ros-1 protein**. No language limits were applied. Take PubMed as an example, the retrieval strategy was as follows: #1 Search” Carcinoma, Non-Small-Cell Lung”[Mesh] OR (((((((((((((Carcinomas, Non-Small-Cell Lung)[Title/ Abstract]) OR Lung Carcinoma, Non-Small-Cell)[Title/ Abstract]) OR Lung Carcinomas, Non-Small-Cell)[Title/ Abstract]) OR Non-Small-Cell Lung Carcinomas[Title/ Abstract]) OR Non-Small Cell Lung Cancer[Title/ Abstract]) OR Non-Small-Cell Lung Carcinoma[Title/ Abstract]) OR Non Small Cell Lung Carcinoma[Title/ Abstract]) OR Carcinoma, Non-Small Cell Lung[Title/Abstract]) OR Non-Small Cell Lung Cancer[Title/Abstract]) OR non-small cell lung cancer[Title/Abstract]) OR NSCLC[Title/Abstract]) OR lung neoplasm[Title/Abstract]) OR lung cancer[Title/Abstract]) OR lung carcinoma[Title/Abstract]) OR pulmonary neoplasm) OR pulmonary cancer) OR pulmonary carcinoma), #2 Search ((((**ROS1**[Title/Abstract]) OR **ROS-1**[Title/Abstract]) OR c-ros protein[Title/Abstract]) OR **Ros1** proto-oncogene protein[Title/Abstract]) OR Ros-1 protein[Title/Abstract], #3 (“2015/01/01”[Date-Publication]: “2019/10/27”[Date -Publication]), #4 Search #1 AND #2 AND #3.

**Study selection**

Two independent investigators separately screened the study selection against the inclusion criteria. The differences between the two investigators were settled by negotiation. If necessary, we try to contact the author of the original article for accurate information. The inclusion criteria for articles were as follows: (I) involved NSCLC patients, regardless of pathology type; (II) **ROS1** gene rearrangement was involved; (III) included the clinical characteristics of patients with **ROS1** gene rearrangement, e.g., gender, smoking status, pathological type and clinical stage; and (IV) included the number of **ROS1** gene rearranged patients and
their specific clinical characteristics. Articles that met any of the following criteria were excluded: (I) case reports, reviews, guidelines, or correspondence; (II) containing no highly relevant information, non-original articles; (III) lacking sufficient data to calculate odds ratios (ORs) and 95% confidence intervals (CIs).

**Data collection and quality assessment**

Data from eligible studies were collected by two independent researchers. During the data extraction period, in order to reduce the offset and error probability, the two researchers exchanged the original data tables for checking. If a dispute arises, the third researcher shall conduct arbitration. The information collected is as follows: the first author's name, year of publication, country of the study, total number of patients, race, age (range), detection technique for the ROS1 gene rearrangement, number of ROS1-rearranged patients, and tumor pathological type.

The quality assessment of the studies included was assessed using an 11 items checklist which was recommended by Agency for Healthcare Research and Quality (AHRQ). If an item was selected “NO” or “UNCLEAR” as answer, the score for “0”; If an item was selected “YES” as answer, then the items’ score for “1”. Article quality was assessed as follows: low quality =0–3; moderate quality =4–7; high quality =8–11. This group into the article 10 articles belongs to high quality articles, article 29 belongs to medium quality articles. From the results of quality evaluation, the original research is reliable. The data sources and inclusion exclusion criteria of all the included cross-sectional studies are explicit, which makes the confounding factors under reasonable control. In addition, some studies provided the response rate and follow-up results. The specific quality score results are shown in Table 1.

**Statistical analysis**

The relationship between ROS1 gene rearrangement and four clinical characteristics was studied in NSCLC patients. The four clinical characteristics were: gender, smoking status, pathological type, and clinical stage. Heterogeneity was tested by I² statistic and χ²-based Q test (56). With I² higher than 50% or P value of Q statistic's test under 0.10 showing significant heterogeneity significant (57). If the heterogeneity was reduced to a mild degree after the source had been identified, then a fixed effects model was applied. If the heterogeneity was still obvious, then a random effects model was adopted to pool effect sizes of ORs and their corresponding 95% CIs (58). Because I² was not an absolute measure of heterogeneity (59), sensitivity analysis was also carried out to rule out the possibility that heterogeneity among studies could be offset. We observed whether the OR value of each study was located near the aggregate OR value of the total effect quantity and decided whether to delete some articles. Quantitative evaluation of publication bias by Begg's funnel plot and Egger's test, P<0.05 shows that publication bias has statistical significance. If the number of studies included is less than 10, the stability of the conclusions will be further evaluated after the publication bias is eliminated by trim-fill method.

**Results**

**Study selection and characteristics**

The flow chart of the literature selection process is exhibited in Figure 1. Our systematic search generated 2,619 studies, of which 2,580 studies were duplicates, case reports, reviews, guidelines, or correspondence, etc. Finally, 39 studies involving 25,055. This meta-analysis included 39 studies that described the clinical characteristics of patients with ROS1 gene rearrangement. Due to the high expression of ROS1, in alveolar type II lung cells, macrophages, and osteoclasts during bone metastasis, which may lead to an increase in false positive rate, when used alone in the immunohistochemical (IHC) method to determine the results of ROS1 gene rearrangement (60). Only the results of ROS1 gene rearrangement by immunohistochemical method were excluded. The specific screening results are shown in Table 1.

**Relationship between ROS1 gene rearrangement and gender in NSCLC patients**

Thirty-five studies in the analysis discussed the correlation between ROS1-rearranged patients and gender. These studies involved a total of 8,440 male patients and 12,506
Table 1 Characteristics of included studies

| Study/first author | Year | Country | Race     | No. of patients | Age [range] | ROS1(+) | Detection methods | Tumor type | Quality score |
|--------------------|------|---------|----------|----------------|-------------|---------|-------------------|------------|---------------|
| Bai et al. (17)    | 2019 | China   | Asian    | 379            | 60 [19–89]  | 12      | PCR               | NSCLC      | 6             |
| Bu et al. (18)     | 2017 | China   | Asian    | 1875           | NA          | 11      | FISH              | NSCLC      | 8             |
| Chen et al. (19)   | 2019 | China   | Asian    | 89             | 32 [18–35]  | 1       | RT-PCR            | Ade        | 8             |
| Chen et al. (20)   | 2019 | China   | Asian    | 205            | 57 [38–82]  | 5       | RT-PCR            | NSCLC      | 6             |
| Clavé et al. (21)  | 2016 | USA     | Caucasian| 314            | 64 [25–91]  | 5       | FISH              | Ade        | 6             |
| Dugay et al. (22)  | 2017 | France  | Caucasian| 71             | 65.6±10.7   | 15      | FISH              | NSCLC      | 6             |
| Fu et al. (23)     | 2015 | China   | Asian    | 204            | NA          | 4       | FISH              | Ade        | 8             |
| Gao et al. (24)    | 2018 | China   | Asian    | 356            | NA          | 15      | FISH/PCR          | NSCLC      | 6             |
| Ha et al. (25)     | 2015 | Korea   | Asian    | 198            | 60 [29–81]  | 2       | PCR               | Ade        | 6             |
| He et al. (26)     | 2019 | China   | Asian    | 347            | 62 [26–86]  | 6       | RT-PCR            | NSCLC      | 6             |
| Huang et al. (27)  | 2019 | China   | Asian    | 300            | NA          | 5       | NGS               | NSCLC      | 6             |
| Jiang et al. (28)  | 2017 | China   | Asian    | 297            | NA          | 3       | IHC/FISH          | NSCLC      | 6             |
| Joshi et al. (29)  | 2019 | India   | Asian    | 535            | 60 [16–84]  | 22      | FISH              | NSCLC      | 8             |
| Jürmeister et al. (30) | 2015 | Germany | Caucasian| 473            | 54 [29–75]  | 4       | FISH              | NSCLC      | 6             |
| Kuang et al. (31)  | 2019 | China   | Asian    | 113            | NA          | 4       | NGS               | NSCLC      | 6             |
| Lee et al. (32)    | 2018 | Korea   | Asian    | 407            | 66.9±12.1   | 14      | IHC/FISH          | NSCLC      | 6             |
| Li et al. (33)     | 2017 | China   | Asian    | 76             | 61 [33–85]  | 1       | PCR               | NSCLC      | 6             |
| Li et al. (34)     | 2018 | China   | Asian    | 116            | 51 [28–86]  | 5       | RT-PCR            | NSCLC      | 6             |
| Li et al. (35)     | 2018 | China   | Asian    | 162            | NA          | 20      | RT-PCR            | NSCLC      | 8             |
| Liang et al. (36)  | 2016 | China   | Asian    | 332            | 59 [32–77]  | 13      | FISH              | Ade        | 6             |
| Liu et al. (37)    | 2015 | China   | Asian    | 369            | 59 [27–84]  | 16      | RT-PCR            | Ade        | 6             |
| Liu et al. (38)    | 2018 | China   | Asian    | 139            | NA          | 1       | RT-PCR            | NSCLC      | 6             |
| Mazieres et al. (39)| 2019 | France  | Mixed    | 551            | 60 [29–83]  | 7       | PCR/NGS           | NSCLC      | 8             |
| Plodkowski et al. (40)| 2015 | USA     | Caucasian| 73             | 61 [36–89]  | 17      | FISH              | Ade        | 6             |
| Qiao et al. (41)   | 2017 | China   | Asian    | 514            | 60 [30–85]  | 5       | PCR               | NSCLC      | 6             |
| Rogers et al. (42) | 2015 | Australia| Caucasian| 362            | 73 [36–97]  | 3       | FISH              | Ade        | 6             |
| Song et al. (43)   | 2017 | China   | Asian    | 385            | 58          | 58      | FISH              | NSCLC      | 8             |
| Vlajnic et al. (44) | 2018 | Australia| Caucasian| 295           | 70 [30–100] | 13      | FISH              | NSCLC      | 8             |
| Wang et al. (45)   | 2015 | China   | Asian    | 41             | 26.4±3.5    | 2       | PCR               | NSCLC      | 8             |
| Wang et al. (46)   | 2017 | China   | Asian    | 172            | 59 [26–80]  | 3       | FISH PCR           | Ade        | 6             |
| Wu et al. (47)     | 2015 | China   | Asian    | 127            | 61 [26–85]  | 5       | FISH              | Ade        | 8             |
| Wu et al. (48)     | 2016 | China   | Asian    | 238            | 61 [27–85]  | 10      | FISH              | NSCLC      | 6             |
| Yang et al. (49)   | 2017 | China   | Asian    | 86             | 55 [33–85]  | 2       | RT-PCR            | NSCLC      | 6             |
| Yang et al. (50)   | 2019 | China   | Asian    | 3,487          | 55 [24–85]  | 55      | RT-PCR            | NSCLC      | 6             |
| Zhang et al. (51)  | 2016 | China   | Asian    | 280            | 60 [31–87]  | 5       | PCR               | Ade        | 6             |
| Zhang et al. (52)  | 2018 | China   | Asian    | 3,345          | NA          | 90      | RT-PCR            | NSCLC      | 8             |
| Zhang et al. (53)  | 2019 | China   | Asian    | 6,066          | 60.1±10.6   | 157     | RT-PCR            | NSCLC      | 6             |
| Song et al. (54)   | 2017 | China   | Asian    | 732            | 60 [28–81]  | 32      | RT-PCR            | Ade        | 8             |
| Zhong et al. (55)  | 2015 | China   | Asian    | 302            | 60 [19–82]  | 12      | PCR               | NSCLC      | 6             |

ROS1(+), ROS1 rearrangement positive; RT-PCR, reverse transcriptase PCR; FISH, fluorescence in situ hybridization; IHC, immunohistochemistry; NSCLC, non-small cell lung cancer; NGS, next generation sequencing; Ade, adenocarcinoma.
female patients, and \textit{ROS1} gene rearrangement was observed in 288 females (3.41\%, 288/8,440) and 217 males (1.73\%, 217/12,506) patients. We found no significant heterogeneity (\(I^2=0\%\), \(P=0.48\); Figure S1A), we didn't find obvious heterogeneity by sensitivity analysis (Figure S1B), and so the fixed effects model was selected to incorporate the effect size. Pooled results suggested that female NSCLC patients displayed a higher rate of \textit{ROS1} gene rearrangement when compared with male patients (OR =1.94, 95\% CI: 1.62–2.32\%, \(P<0.05\); Figure S1A). Publication bias was tested by Begg's test (\(P=0.532\)) and Egger's test (\(P=0.574\)), and no publication bias was observed.

\textbf{Relationship between \textit{ROS1} gene rearrangement and smoking status in NSCLC patients}

Thirty studies in the analysis explored the relationship between \textit{ROS1} gene rearrangement and smoking status. A total of 10,161 non-smokers and 8,690 smokers or former

\begin{figure}
\centering
\includegraphics[width=	extwidth]{flowchart.png}
\caption{Flow chart of the study selection for the meta-analysis.}
\end{figure}
smokers were calculated, and \( \text{ROS1} \) gene rearrangement was observed in 328 non-smokers (3.22%, 328/10,161) and 97 smokers or former smokers (1.11%, 97/8,690). No significant heterogeneity was observed (\( I^2 = 8\% \), \( P = 0.34 \); Figure S2A). Sensitivity analysis was performed (Figure S2B) and no significant heterogeneity was discovered, so the fixed effects model was selected to incorporate the effect size. Pooled results suggested that patients with no smoking history had an increased rate of \( \text{ROS1} \) gene rearrangement compared with smokers (OR = 2.82, 95% CI: 2.24–3.55%, \( P < 0.05 \) (Figure S2A). Publication bias was checked by Begg’s test (\( P = 0.392 \)) and Egger’s test (\( P = 0.321 \)), and no publication bias was found to exist.

**Relationship between \( \text{ROS1} \) gene rearrangement and pathological subtype in NSCLC patients**

There were 29 studies about the relationship between \( \text{ROS1} \) gene rearrangement and pathological subtype. Mild heterogeneity was observed (\( I^2 = 18\% \) and \( P = 0.2 \), Figure S3A). Sensitivity analysis showed that 2 articles [Yang et al. (2019) (50) and Zhang et al. (2018) (52)] had significantly influenced the final results (Figure S3B). Therefore, these studies were omitted from the data analysis. Among the 27 remaining articles, the heterogeneity decreased obviously (\( I^2 = 0\% \) and \( P = 0.82 \); Figure S3C). A total number of 12,256 adenocarcinoma patients and 3,567 non-adenocarcinoma patients were included in the calculation, with \( \text{ROS1} \) gene rearrangement observed in 281 adenocarcinoma patients (2.30%, 281/12,256) and 49 non-adenocarcinoma patients (1.37%, 49/3,567). We did a sensitivity analysis again and found no other studies that had a significant impact on heterogeneity (Figure S3D), so the fixed effects model was selected to incorporate the effect size. Pooled results suggested that adenocarcinoma patients had a higher rate of \( \text{ROS1} \) gene rearrangement than non-adenocarcinoma patients (OR = 1.55, 95% CI: 1.14–2.11%, \( P < 0.05 \); Figure S3C). Publication bias was tested by Begg’s test (\( P = 0.802 \)) and Egger’s test (\( P = 0.798 \)) and no publication bias was detected.

**Relationship between \( \text{ROS1} \) gene rearrangement and clinical stage in NSCLC patients**

There were 20 studies on the relationship between \( \text{ROS1} \) gene rearrangement in NSCLC patients and clinical stage. A total of 8,683 patients with stage III–IV and 5,904 patients with stage I–II were calculated. The merged frequency of \( \text{ROS1} \) gene rearrangement was 2.81% (244/8,683) in patients with stage III–IV disease and 1.61% (95/5,904) in patients with stage I–II disease. No significant heterogeneity was found (\( I^2 = 0\% \), \( P = 0.568 \); Figure S4A). Sensitivity analysis was performed (Figure S4B) and no significant heterogeneity was discovered, so the fixed effects model was selected to incorporate the effect size. The merged results showed that an advanced clinical stage (III–IV) was associated with a significantly higher \( \text{ROS1} \) rearrangement rate (OR = 1.50, 95% CI: 1.15–1.94%, \( P = 0.003 \); Figure S4A). Begg’s test (\( P = 0.456 \)) and Egger’s test (\( P = 0.917 \)) publication bias.

**Prevalence of \( \text{ROS1} \) gene rearrangement in lung adenocarcinoma**

There were 29 studies about the relationship between \( \text{ROS1} \) gene rearrangement and pathological subtype. Mild heterogeneity was observed (\( I^2 = 18\% \) and \( P = 0.2 \), Figure S3A). Sensitivity analysis showed that 2 articles [Yang et al. (2019) (50) and Zhang et al. (2018) (52)] had significantly influenced the final results (Figure S3B). Therefore, these studies were omitted from the data analysis. Among the 27 remaining articles, the heterogeneity decreased obviously (\( I^2 = 0\% \) and \( P = 0.82 \); Figure S3C). A total number of 12,256 adenocarcinoma patients and 3,567 non-adenocarcinoma patients were included in the calculation, with \( \text{ROS1} \) gene rearrangement observed in 281 adenocarcinoma patients (2.30%, 281/12,256) and 49 non-adenocarcinoma patients (1.37%, 49/3,567). We did a sensitivity analysis again and found no other studies that had a significant impact on heterogeneity (Figure S3D), so the fixed effects model was selected to incorporate the effect size. Pooled results suggested that adenocarcinoma patients had a higher rate of \( \text{ROS1} \) gene rearrangement than non-adenocarcinoma patients (OR = 1.55, 95% CI: 1.14–2.11%, \( P < 0.05 \); Figure S3C). Publication bias was tested by Begg’s test (\( P = 0.802 \)) and Egger’s test (\( P = 0.798 \)) and no publication bias was detected.

**Discussion**

This meta-analysis of these studies showed that \( \text{ROS1} \) gene rearrangement was more common in female patients with no smoking history, adenocarcinoma patients, and patients at a more advanced stage (stage III–IV). The study also showed that prevalence rate of adenocarcinoma of \( \text{ROS1} \) gene rearrangement was 2.49% (95% CI: 1.92–3.11%; Figure S5A) in adenocarcinoma.

There were 29 studies involved in the analysis of the prevalence of \( \text{ROS1} \) gene rearrangement. A total of 19,385 adenocarcinoma patients were analyzed and 503 were positive for the \( \text{ROS1} \) gene rearrangement. Significant heterogeneity was detected (\( I^2 = 78.15\% \), \( P = 0 \); Figure S5A). Sensitivity analysis was carried out to find the source of heterogeneity, none of the literature caused great interference to the results of this meta-analysis, indicating that the stability of this study is acceptable, as shown in the picture (Figure S5B). The lack of control groups in single-arm studies could serve as one reason for the increased heterogeneity. Therefore, the random effects model was used to obtain the total effect quantity, as shown in the forest plot above. Pooled analysis of all 29 studies showed a prevalence of 2.49% (95% CI: 1.92–3.11%; Figure S5A) in adenocarcinoma.
comparison, the proportion of ROS1-rearranged non-
adenocarcinoma patients in this study was 1.37%. The
crison rearrangement is higher in adenosquamous carcinoma and pleomorphic carcinoma
than adenocarcinoma, which is one of the reasons for the higher rate (62).

Although similar articles have emerged recently. Kho et al. has made a similar analysis (63), this article was published in a short format in the Ann Oncol supplement and only focused on Asian population. We have analyzed more data collected from all over the world in our meta-analysis, the summarized results have little heterogeneity and good data stability. We did sensitivity analysis and our summarized results have little heterogeneity, indicating good data stability. In addition, we used the AHRQ scale to evaluate the data quality from the original articles and only kept the satisfactory data set in our meta-analysis.

The relationship between NSCLC patients with ROS1 gene rearrangement and treatment needs further prospective study. The treatment of patients with ROS1 gene rearrangement is classical chemotherapy and/or targeted therapy, and the commonly used targeted drug is crizotinib. Scheffler's study showed that 64.2% of patients with ROS1 gene rearrangement responded to chemotherapy, and the average overall survival (OS) of patients who received chemotherapy and crizotinib was estimated at 5.3 years (64). The results are somewhat consistent with the recently reported efficacy of crizotinib in the treatment of advanced NSCLC (65,66), recent studies have shown a change in the choice of treatment for patients with advanced ROS1 rearrangement in NSCLC. Patients with first-line selection of platinum-pemetrexed and crizotinib had different objective remission rates (85.1% vs. 86.7%) and progression-free survival (8.6 vs. 18.4 months), and there was no significant difference in OS (less than vs. 28.4 months) between the two groups (66). The relationship between ROS1 gene rearrangement and drug therapy needs further research to provide evidence.

There are some limitations in the current meta-analysis. Firstly, there may be potential publication bias in our analysis. Although we conducted a systematic search of many databases and berg’s tests didn’t show significant publication bias, we still can’t rule out the possibility of omitting relevant research. Secondly, the purpose of our study of the ROS1 gene rearrangement was to provide guidance for the selection of clinical treatment regime. The selection of targeted therapy, chemotherapy and targeted combination chemotherapy is related to overall survival.

There was no standardized scheme among the 39 studies in this meta-analysis. The selection of drugs for patients with ROS1 gene rearrangement is still under investigation, and more large-scale clinical trials or prospective studies are required. Finally, ROS1 differs from ALK, in which the kinase is activated by the rearrangement partner providing a dimeric domain to induce constitutive oligomers, while the constitutive activation mechanism of ROS1 fusion proteins is unclear. Indeed, many known ROS1 fusion partners do not contain dimer domains (67). Attention should be paid to the clinical effect of TKI in patients with ROS1 gene rearrangement, and further related observation and research are needed.

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Footnote

Reporting Checklist: The authors have completed the PRISMA reporting checklist. Available at http://dx.doi.org/10.21037/tcr-20-1813

Conflicts of Interest: All authors have completed the ICMJE uniform disclosure form (available at http://dx.doi.org/10.21037/tcr-20-1813). The authors have no conflicts of interest to declare.

Ethical Statement: The authors are accountable for all aspects of the work in ensuring that questions related to the accuracy or integrity of any part of the work are appropriately investigated and resolved.

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Figure S1 The relationship between \textit{ROS1} gene rearrangement and gender in NSCLC patients. (A) Forest plot of odds ratios (ORs) and 95\% confidence intervals (CIs) from the related studies. The first author has the same last name. (B) Sensitivity analysis of the relationship between \textit{ROS1} gene rearrangement and gender in NSCLC patients. The order of first names of the same surname is shown in Table 1.
Figure S2 The relationship between ROS1 gene rearrangement and smoking status in NSCLC patients. (A) Forest plot of odds ratios (ORs) and 95% confidence intervals (CIs) from the related studies. The first author has the same last name. (B) Sensitivity analysis of the relationship between ROS1 gene rearrangement and smoking status in NSCLC patients. The order of first names of the same surname is shown in Table 1.
Figure S3 Relation between ROS1 gene rearrangement positive and Pathological subtype in NSCLC patients. (A) Forest plot of odds ratios (ORs) and 95% confidence intervals (CI) from related study. The first author has the same last name. (B) Sensitivity analysis of the relationship between ROS1 gene rearrangement positive and pathological subtype in NSCLC patients. The order of first names of the same surname is shown in Table 1. (C) The relationship between ROS1 gene rearrangement and pathological subtype in NSCLC patients. Forest plot of odds ratios (ORs) and 95% confidence intervals (CIs) from the related studies. The first author has the same last name. (D) Sensitivity analysis of the relationship between ROS1 gene rearrangement and pathological subtype in NSCLC patients. The order of first names of the same surname is shown in Table 1.
**Figure S4** The relationship between *ROS1* gene rearrangement and clinical stage in NSCLC patients. (A) Forest plot of odds ratios (ORs) and 95% confidence intervals (CIs) from the related studies. The first author has the same last name. (B) Sensitivity analysis of the relationship between *ROS1* gene rearrangement and clinical stage in NSCLC patients. The order of first names of the same surname is shown in Table 1.
Figure S5 Prevalence of *ROS1* gene rearrangement in NSCLC patients with adenocarcinoma. (A) Forest plot of odds ratios (ORs) and 95% confidence intervals (CIs) from each study. The first author has the same last name. (B) Sensitivity analysis of prevalence of *ROS1* gene rearrangement in lung adenocarcinoma. The order of first names of the same surname is shown in Table 1.