The prognosis for patients with newly diagnosed glioblastoma receiving bevacizumab combination therapy: a meta-analysis

Ke-Li Liao¹
Song Huang¹
Yu-Peng Wu²

¹Department of Neurosurgery, Zigong First People’s Hospital, Zigong, Sichuan, People’s Republic of China; ²Department of Neurosurgery, The Second Hospital of Hebei Medical University, Shijiazhuang, Hebei, People’s Republic of China

Background: A combination of temozolomide (TMZ) and radiotherapy and subsequent adjuvant chemotherapy is the gold standard of treatment for glioblastoma (GB). Bevacizumab (BEV), a humanized monoclonal antibody that blocks the effects of vascular endothelial growth factor A, has produced impressive response rates for recurrent GB and has been approved as second-line therapy. The efficacy and safety of BEV in newly diagnosed GB are not known.

Aim: This systematic meta-analysis was undertaken to evaluate the value of combination therapy involving BEV in newly diagnosed GB.

Methods: Electronic databases were searched for eligible literature up to October 2017. Randomized controlled trials assessing the efficacy and safety of BEV in patients with newly diagnosed GB were included, of which the main outcomes were progression-free survival (PFS), overall survival (OS), and adverse events (AEs). All the data were pooled with the corresponding 95% confidence intervals (CIs) using RevMan software. Sensitivity analyses and heterogeneity were quantitatively evaluated.

Results: A total of six randomized controlled trials were included in this analysis. The experimental BEV group had significantly improved the overall PFS (OR = 0.46, 95% CI = 0.26–0.81, \(P = 0.007\)), as well as PFS at 6 months (OR = 3.47, 95% CI = 2.85–4.22, \(P < 0.00001\)) and PFS at 12 months (OR = 2.02, 95% CI = 1.66–2.46, \(P < 0.00001\)), respectively. However, there were no significant differences in PFS at 24 months with BEV (OR = 0.95, 95% CI = 0.61–1.48, \(P = 0.82\)). OS at 6 months (\(P = 0.07\)) and 24 months (\(P = 0.07\)) was not significantly improved with BEV in patients with newly diagnosed GB. However, the meta-analysis on the OS at 12 months showed differences with BEV (OR = 1.24, 95% CI = 1.03–1.50, \(P = 0.02\)).

Conclusion: Our study indicates that addition of BEV for newly diagnosed GB resulted in a superior PFS rate. However, the combination therapy involving BEV did not improve OS. Future investigations are needed to analyze whether BEV helps improve OS efficacy.

Keywords: bevacizumab, glioblastoma, newly diagnosed, meta-analysis, neoadjuvant

Introduction

Glioblastoma (GB) is the most aggressive brain malignancy in adults.¹ Even with the advances in imaging and available standard treatment, the survival and quality of life (QoL) of patients with GB still remain poor.²,³ When GB is diagnosed, the median overall survival (OS) with a combination of conventional radiation and temozolomide (TMZ) is only 14 months after diagnosis.³ Therefore, it is necessary to search for more effective treatment options for newly diagnosed GB.

GB is a highly vascularized tumor with upregulated vascular endothelial growth factor (VEGF), which inhibits tumor growth and progression.⁴,⁵ As an antiangiogenic
(anti-VEGF) agent, bevacizumab (BEV) serves as a potentially therapeutic option for GB. Based on the promising outcomes of two Phase II trials, the US Food and Drug Administration (FDA) has approved BEV in the second-line treatment of recurrent GB in 2009. The addition of BEV produces impressive progression-free survival (PFS) and response rate compared to standard treatment for recurrent GB, while the value of BEV on OS is still a matter of debate.

At the time of the approval of BEV in recurrent disease, several randomized controlled trials (RCTs) were launched to investigate the value of BEV in addition to the treatment for newly diagnosed GB. Treatment combined with BEV exhibited significant activity in PFS for newly diagnosed GB patients, while further evaluation is needed to assess disease progression after antiangiogenic therapy.

In order to make care for newly diagnosed GB more rational, we conducted a meta-analysis of RCTs to evaluate the therapeutic value of BEV compared with standard therapy (ST).

**Methods**

**Search strategy**

Two investigators independently searched the electronic databases PubMed, Embase, and Cochrane library up to October 2017. We searched for all randomized clinical trials evaluating the value of BEV in patients with newly diagnosed GB. The process involved finding all articles with the keywords “bevacizumab” AND “glioblastoma” AND “newly diagnosed” AND “efficacy”, and the relevant Medical Subject Heading (MeSH) terms were searched. The reference lists of all articles that dealt with the topic of interest were also hand-searched to check for additional relevant publications. The search was restricted to trials published in the English language.

**Eligibility criteria**

Studies that met the following criteria were included in the meta-analysis: 1) the studies were designed as RCTs; 2) studies focused on newly diagnosed GB treated with BEV; 3) the outcomes of interest regarding the treatment efficacy (survival), safety (adverse events [AEs]), and hazard ratios (HRs) with corresponding 95% CIs were provided; 4) only the full texts were included. The studies that did not meet the above inclusion criteria were excluded from the meta-analysis.

**Quality assessment**

Study quality was assessed using the Jadad seven-item scale recommended by The Cochrane Handbook for Systematic Reviews of Interventions. The Jadad seven-item scale was introduced to evaluate the overall methodological quality of RCTs.

**Data extraction**

The data extraction was conducted independently by two authors. Disagreement was resolved by consensus. The main characteristics extracted from the selected studies were the following: first author family name, publication year, trial name, study design, sample size, and the outcomes of interest (AEs, PFS, and OS). The corresponding hazard ratios (HRs) with 95% confidence intervals (CIs) were used to describe the main outcomes of the studies, including OS, PFS, and AE data, and 95% CIs were calculated for each estimate.

**Statistical analysis**

The main outcomes of the studies were OS, PFS, and AEs. If HRs and corresponding 95% CIs were reported, the lnHRs and the corresponding ln lower limits and ln upper limits were used as data points in pooling analysis. While, if the study did not provide HRs or 95% CIs, the only available data were in the form of Kaplan–Meier (K–M) curves. Survival data were extracted from the form of the K–M survival curve, according to the methods described by Tierney et al.

The I² statistical test was used to further examine statistical heterogeneity between the trials. Studies with an I²≥50% were considered to exhibit moderate and high heterogeneity, and those with I²<50% were considered to have low heterogeneity. Summary HRs were calculated by using fixed-effect models when there was low heterogeneity among studies. Otherwise, random-effect models were used. A P-value <0.05 was considered to be statistically significant. All analyses were conducted with Review Manager Version 5.3 software (Revman; The Cochrane Collaboration, Oxford, UK). Findings of our meta-analysis are shown in forest plots. The Begg’s test and the Egger’s test were used to evaluate publication bias.

**Results**

**Search results and characteristics of studies**

A total of 214 studies were retrieved initially for evaluation. Based on the criteria described in the “Methods” section, 10 publications were evaluated in more detail, but some did not provide enough detail of the outcomes of two approaches. Therefore, we had a final total of six RCTs assessing the value of BEV in patients with newly diagnosed GB. The search process is described in Figure 1.

All included papers in this study were based on moderate-to high-quality evidence. Table 1 describes the primary characteristics of the eligible studies in more detail.
The prognosis for patients with newly diagnosed GB treated with added BEV

Clinical and methodological heterogeneity
Pooled analysis of PFS comparing the addition of BEV with the control group

Pooling the PFS data from five studies showed that BEV prolonged the PFS (HR =0.69, 95% CI =0.63–0.77, P<0.00001) compared with the control group (Figure 2).

Subgroup analysis of PFS comparing the addition of BEV with the control group
Overall, the six studies that reported data on PFS at different months are shown in Figures 3–5. Pooled data showed that the PFS data achieved advantage with BEV agents, with the pooled OR being 3.47 (95% CI 2.85–4.22, P<0.00001) at 6 months (Figure 3); 2.02 (95% CI 1.66–2.46, P<0.00001) at

Table 1 The primary characteristics of the eligible studies in more detail

| Study                  | Trials      | Mean age, years | Cases, n | Treatment                      | Follow-up (months) |
|------------------------|-------------|-----------------|----------|-------------------------------|-------------------|
|                        |             | Experiment      | Control  | Experiment                     |                   |
| Gilbert et al, 2014    | AVAglio     | NA              | NA       | 312                           | BEV+RT+TMZ        | 30                |
| Chinot et al, 2014     | RTOG 082512 | 57              | 56       | 458                           | BEV+RT+TMZ        | 32                |
| Chauffert et al, 2014  | TEMAVIR     | 60.2            | 60.9     | 60                            | BEV+RT+TMZ        | 24                |
| Herrlinger et al, 2016 | GLARIUS     | 56              | 56       | 116                           | BEV+RT+TMZ+IRI    | 36                |
| Carlson et al, 2015    | NA          | 55.9            | 59.5     | 30                            | IMRT+BEV+TMZ      | 60                |
| Balana et al, 2016     | GENOM 009   | 62.9            | 62       | 48                            | BEV+RT+TMZ        | NA                |

Abbreviations: BEV, bevacizumab; IMRT, intensity-modulated radiotherapy; IRI, irinotecan; NA, not available; RT, radiotherapy; TMZ, temozolomide.
Due to the limited data in all studies, systematic evaluations of AE data were not possible in this meta-analysis. Gilbert et al report that toxicities, hypertension, thromboembolic events, intestinal perforation, and neutropenia were observed in the BEV group. Over time, patients treated with BEV have an increased symptom burden following a worse QoL, as well as debilitating neurological symptoms. In Chinot et al, grade 3 or higher AEs (66.8% vs 51.3%) were more frequent in the BEV group.

**Discussion**

GB is the most frequent malignant brain tumor with a poor prognosis. Patients with recurrent GBs have a poor OS, and available therapies have a limited impact on prognosis. Therefore, development of a new approach is essential to improve the outcomes in patients with newly diagnosed GB. In Phase III RCTs, the addition of TMZ to chemotherapy (CRT) and subsequent adjuvant chemotherapy has been the standard therapy with newly diagnosed GB and is thought to be the backbone for further understanding therapy choices. Research reports have shown that GB has multistep cytostatic effects that alter neovascularization of brain tissue to form new blood vessels, which may help slow tumor progression and proliferation. VEGF is overexpressed in malignant gliomas and has been used as a therapeutic target for brain tumors.

BEV, a humanized monoclonal antibody against the VEGF ligand, has received FDA approval for recurrent GBs, and proliferation.

**Subgroup analysis of OS comparing the addition of BEV with the control group**

Subgroup analysis of OS data at different months was available for all six RCTs. Analysis showed that the results of OS at 6 months (HR =1.28, 95% CI =0.98–1.67, P=0.07) (Figure 7) with BEV were no longer significant in patients with newly diagnosed GB, while the meta-analysis on the OS at 12 months showed differences with BEV (HR =1.24, 95% CI =1.03–1.50, P=0.02) (Figure 8). However, OS at 24 months still did not reach statistically significant difference (HR =1.22, 95% CI =0.98–1.52, P=0.07) (Figure 9).

**Pooled analysis of AEs comparing the addition of BEV with the control group**

Due to the limited data in all studies, systematic evaluations of AE data were not possible in this meta-analysis.
GBMs in 2009. The results have suggested that a potential benefit could be achieved by BEV for newly diagnosed GB. To assess whether BEV would be safe and effective for the treatment of newly diagnosed GB, we conducted a meta-analysis to evaluate the value of therapy with the combination of BEV.

All included studies chose OS and PFS as the primary end points. At the time of disease progression, crossover regimen may continue to benefit patients following an initial response to therapy with or without BEV. Therefore, in newly diagnosed GBs, the end point of PFS has beneficial effects in evaluating the potential effects of combined treatment with BEV than OS.

Data from the trial of Carlson et al. showed that compared with placebo plus TMZ and radiotherapy, the combination using BEV did not improve OS, but increases PFS, although it is statistically nonsignificant (P=0.39). In the AVAglio and RTOG 0825 studies, the PFS was significantly improved with the addition of BEV, but OS did not show benefit.

In this study, we conclude that the combination of BEV for newly diagnosed GB is beneficial in terms of prolonging median PFS but not OS. Our results did not indicate any benefit from BEV for newly diagnosed GB in terms of median OS. The AVAglio, RTOG 082512, and GLARIUS trials obtained similar results. In our analysis, the pooled analysis did not show that the PFS benefit translates into OS prolongation. As potential reasons for this observation, patients with GB exhibited worse neurocognition and poor survival rate with prolonged use of BEV, which might be caused by BEV resistance. Resistance to chemotherapy was considered to influence the effectiveness of BEV treatment for GB. As an antiangiogenic (anti-VEGF) agent, BEV has been investigated as complementary to standard chemotherapy to suppress tumor growth. Due to the different angiogenesis pathways of the VEGF genes, there may be benefit in continuing BEV treatment even after resistance to chemotherapy. However, the potential antitumor effects underlying resistance to antiangiogenic agents are yet to be fully evaluated. The point of molecular signatures may reveal subsets of GBs that are particularly sensitive or resistant to BEV. In additional analyses of subgroups of patients based on different genetic mutations, we may identify patients who had a selected survival benefit response to BEV. Furthermore, BEV’s radiographic effect has been reported to be associated with an increased incidence of PFS. BEV stabilizes the blood–brain barrier, minimizing the ability of the magnetic resonance imaging (MRI) contrast agent gadolinium to reach the tumor, thus

**Figure 4** Pooled analysis of PFS at 12 months comparing the bevacizumab-addition group with the control group.

| Study or subgroup | Experimental Events Total | Control Events Total | Weight (%) | OR M–H, fixed, 95% CI |
|------------------|--------------------------|---------------------|------------|----------------------|
| Chauffert et al, 2014 | 18 60 | 10 60 | 5.0 | 2.14 (0.89, 5.14) |
| Balana et al, 2016 | 3 48 | 4 45 | 2.7 | 0.68 (0.14, 3.24) |
| Carlson et al, 2015 | 19 30 | 8 26 | 2.2 | 3.89 (1.27, 11.86) |
| Gilbert et al, 2014 | 133 311 | 96 309 | 39.1 | 1.66 (1.19, 2.30) |
| Chiot et al, 2014 | 189 458 | 110 463 | 45.6 | 2.25 (1.70, 2.99) |
| Herrlinger et al, 2016 | 35 116 | 8 54 | 5.4 | 2.48 (1.06, 5.81) |
| **Total (95% CI)** | **1,023** | **957** | **100** | **2.02 (1.66, 2.46)** |

Total events: 397, 236

Heterogeneity: \( \chi^2 = 5.39, df = 5 (P=0.37); I^2 = 7\%

Test for overall effect: \( Z = 7.00 \quad (P < 0.00001) \)

**Figure 5** Pooled analysis of PFS at 24 months comparing the bevacizumab-addition group with the control group.

| Study or subgroup | Experimental Events Total | Control Events Total | Weight (%) | OR M–H, fixed, 95% CI |
|------------------|--------------------------|---------------------|------------|----------------------|
| Chauffert et al, 2014 | 3 60 | 1 60 | 2.4 | 3.11 (0.31, 30.73) |
| Balana et al, 2016 | 1 48 | 1 45 | 2.5 | 0.94 (0.06, 15.43) |
| Carlson et al, 2015 | 3 60 | 1 60 | 2.4 | 3.11 (0.31, 30.73) |
| Gilbert et al, 2014 | 17 311 | 27 309 | 63.8 | 0.60 (0.32, 1.13) |
| Chiot et al, 2014 | 13 458 | 8 463 | 19.2 | 1.66 (0.68, 4.05) |
| Herrlinger et al, 2016 | 5 116 | 3 54 | 9.8 | 0.77 (0.18, 3.33) |
| **Total (95% CI)** | **1,053** | **991** | **100** | **0.95 (0.61, 1.48)** |

Total events: 42, 41

Heterogeneity: \( \chi^2 = 5.64, df = 5 (P=0.34); I^2 = 11\%

Test for overall effect: \( Z = 0.23 \quad (P = 0.82) \)
| Study or subgroup | Log(OR) | SE | Weight (%) | OR IV, random, 95% CI | OR IV, random, 95% CI |
|------------------|---------|----|------------|-----------------------|-----------------------|
| Chauffert et al., 2014\(^{11}\) | -0.3567 | 0.2032 | 13.4 | 0.70 (0.47, 1.04) | 0.70 (0.47, 1.04) |
| Balana et al., 2016\(^{10}\) | -0.3857 | 0.2221 | 11.8 | 0.68 (0.44, 1.05) | 0.68 (0.44, 1.05) |
| Gilbert et al., 2014\(^{12}\) | 0.1222 | 0.0994 | 27.5 | 1.13 (0.93, 1.37) | 1.13 (0.93, 1.37) |
| Chint et al., 2014\(^{14}\) | -0.1278 | 0.0749 | 32.1 | 0.88 (0.76, 1.02) | 0.88 (0.76, 1.02) |
| Herrlinger et al., 2016\(^{15}\) | 0.0198 | 0.1848 | 15.2 | 1.02 (0.71, 1.47) | 1.02 (0.71, 1.47) |
| **Total (95% CI)** | 100 | | | 0.91 (0.76, 1.08) | 0.91 (0.76, 1.08) |

Heterogeneity: \(\chi^2=0.02; \chi^2=8.58 \ (df=4 \ (P=0.07); I^2=53\%\)

Test for overall effect: \(Z=1.08 \ (P=0.28)\)

Figure 6 Pooled analysis of OS at 6 months comparing the bevacizumab-addition group with the control group.

**Abbreviations:** CI, confidence interval; IV, inverse-variance; OR, odds ratio; OS, overall survival; SE, standard error.

| Study or subgroup | Experimental Events | Control Events | Total | Weight (%) | OR M–H, fixed, 95% CI | OR M–H, fixed, 95% CI |
|------------------|---------------------|----------------|-------|------------|-----------------------|-----------------------|
| Chauffert et al., 2014\(^{11}\) | 42 | 60 | 129 | 1.00 (0.46, 2.18) | 1.00 (0.46, 2.18) |
| Balana et al., 2016\(^{10}\) | 40 | 48 | 64 | 1.25 (0.44, 3.58) | 1.25 (0.44, 3.58) |
| Carlson et al., 2015\(^{12}\) | 28 | 30 | 58 | 1.17 (0.15, 8.92) | 1.17 (0.15, 8.92) |
| Gilbert et al., 2014\(^{14}\) | 263 | 312 | 575 | 1.14 (0.74, 1.47) | 1.14 (0.74, 1.47) |
| Chint et al., 2014\(^{14}\) | 421 | 458 | 879 | 1.63 (1.06, 2.52) | 1.63 (1.06, 2.52) |
| Herrlinger et al., 2016\(^{15}\) | 109 | 116 | 225 | 0.92 (0.23, 3.69) | 0.92 (0.23, 3.69) |
| **Total (95% CI)** | 1,024 | 957 | 100 | 1.28 (0.98, 1.67) | 1.28 (0.98, 1.67) |

Total events: 903/813

Heterogeneity: \(\chi^2=2.90; \chi^2=5 \ (df=4 \ (P=0.23); I^2=10\%\)

Test for overall effect: \(Z=1.84 \ (P=0.07)\)

Figure 7 Pooled analysis of OS at 6 months comparing the bevacizumab-addition group with the control group.

**Abbreviations:** CI, confidence interval; M–H, Mantel–Haenszel; OR, odds ratio; OS, overall survival.

| Study or subgroup | Experimental Events | Control Events | Total | Weight (%) | OR M–H, fixed, 95% CI | OR M–H, fixed, 95% CI |
|------------------|---------------------|----------------|-------|------------|-----------------------|-----------------------|
| Chauffert et al., 2014\(^{11}\) | 26 | 60 | 86 | 0.87 (0.43, 1.79) | 0.87 (0.43, 1.79) |
| Balana et al., 2016\(^{10}\) | 22 | 48 | 70 | 1.53 (0.67, 3.53) | 1.53 (0.67, 3.53) |
| Carlson et al., 2015\(^{12}\) | 22 | 30 | 52 | 1.22 (0.36, 3.90) | 1.22 (0.36, 3.90) |
| Gilbert et al., 2014\(^{14}\) | 200 | 312 | 512 | 1.09 (0.79, 1.51) | 1.09 (0.79, 1.51) |
| Chint et al., 2014\(^{14}\) | 322 | 458 | 780 | 1.37 (1.04, 1.81) | 1.37 (1.04, 1.81) |
| Herrlinger et al., 2016\(^{15}\) | 98 | 116 | 214 | 1.56 (0.69, 3.51) | 1.56 (0.69, 3.51) |
| **Total (95% CI)** | 1,024 | 957 | 100 | 1.24 (1.03, 1.50) | 1.24 (1.03, 1.50) |

Total events: 690/589

Heterogeneity: \(\chi^2=2.60; \chi^2=5 \ (df=4 \ (P=0.76); I^2=0\%\)

Test for overall effect: \(Z=2.27 \ (P=0.02)\)

Figure 8 Pooled analysis of OS at 12 months comparing the bevacizumab-addition group with the control group.

**Abbreviations:** CI, confidence interval; M–H, Mantel–Haenszel; OR, odds ratio; OS, overall survival.

| Study or subgroup | Experimental Events | Control Events | Total | Weight (%) | OR M–H, fixed, 95% CI | OR M–H, fixed, 95% CI |
|------------------|---------------------|----------------|-------|------------|-----------------------|-----------------------|
| Chauffert et al., 2014\(^{11}\) | 9 | 60 | 69 | 3.35 (0.86, 13.07) | 3.35 (0.86, 13.07) |
| Balana et al., 2016\(^{10}\) | 9 | 48 | 57 | 2.37 (0.67, 8.31) | 2.37 (0.67, 8.31) |
| Carlson et al., 2015\(^{12}\) | 7 | 30 | 37 | 1.01 (0.29, 3.52) | 1.01 (0.29, 3.52) |
| Gilbert et al., 2014\(^{14}\) | 47 | 312 | 359 | 0.92 (0.60, 1.42) | 0.92 (0.60, 1.42) |
| Chint et al., 2014\(^{14}\) | 139 | 458 | 597 | 1.27 (0.95, 1.70) | 1.27 (0.95, 1.70) |
| Herrlinger et al., 2016\(^{15}\) | 28 | 116 | 144 | 1.24 (0.57, 2.73) | 1.24 (0.57, 2.73) |
| **Total (95% CI)** | 1,024 | 957 | 100 | 1.22 (0.98, 1.52) | 1.22 (0.98, 1.52) |

Total events: 239/192

Heterogeneity: \(\chi^2=5.61; \chi^2=5 \ (df=4 \ (P=0.42); I^2=0\%\)

Test for overall effect: \(Z=1.79 \ (P=0.07)\)

Figure 9 Pooled analysis of OS at 24 months comparing the bevacizumab-addition group with the control group.

**Abbreviations:** CI, confidence interval; M–H, Mantel–Haenszel; OR, odds ratio; OS, overall survival.
BEV therapy does not have a place in ST with newly diagnosed GB. We suggest that studies of correlative molecular signatures are needed to identify particular subgroups of patients who will receive benefit from the combination of BEV. Further research is needed to define the best treatment response with the lowest possible toxicity in selecting suitable patients with consideration for their complications and treatment regimen.

Disclosure
The authors report no conflicts of interest in this work.

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