Nowcasting (Short-Term Forecasting) of Influenza Epidemics in Local Settings, Sweden, 2008–2019

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Reliable forecasts of the timing and spatial spread of influenza during seasons and pandemics can meaningfully advance the timing of public health communication campaigns and implementation of resource allocation in healthcare (1). Different types of influenza forecast methods have been developed and applied to support public health response (2). However, although modelers have shown considerable interest in developing infectious disease forecasts, the readiness in the public health community for applying these predictions has been lacking (3). One reason for this discrepancy might be that national public health policies for response to infectious disease outbreaks often assign the responsibility for healthcare resource allocation to local health authorities (i.e., county and municipality governments). For geographic and infrastructural reasons, the timing of the spatial spread of influenza can differ substantially between these administrative units within nations and states. Therefore, a need exists for influenza forecasting methods that harmonize with policy-making responsibilities at local government levels and that are more relevant for public health practitioners.

Another reason for the poor uptake of forecasting methods might be a lack of prospective evaluations of their reliability. To address this issue, the US Centers for Disease Control and Prevention (CDC) has run the Forecast the Influenza Season Collaborative Challenge (FluSight) since the 2013–14 influenza season to prospectively evaluate different methods and data

The timing of influenza case incidence during epidemics can differ between regions within nations and states. We conducted a prospective 10-year evaluation (January 2008–February 2019) of a local influenza nowcasting (short-term forecasting) method in 3 urban counties in Sweden with independent public health administrations by using routine health information system data. Detection-of-epidemic-start (detection), peak timing, and peak intensity were nowcasted. Detection displayed satisfactory performance in 2 of the 3 counties for all nonpandemic influenza seasons and in 6 of 9 seasons for the third county. Peak-timing prediction showed satisfactory performance from the influenza season 2011–12 onward. Peak-intensity prediction also was satisfactory for influenza seasons in 2 of the counties but poor in 1 county. Local influenza nowcasting was satisfactory for seasonal influenza in 2 of 3 counties. The less satisfactory performance in 1 of the study counties might be attributable to population mixing with a neighboring metropolitan area.

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DOI: https://doi.org/10.3201/eid2611.200448
sources for influenza forecasting at the national, regional, and (starting in the 2017–18 influenza season) state level (4). At the local (county and municipality) level, however, few corresponding prospective evaluations based on routine health system data have been reported. Short-term forecasting is denoted as nowcasting (5). Recently, a prospective 5-year appraisal of a local nowcasting method (6) in a county in Sweden (county population ≈460,000) indicated promising results with regard to detection of the local start of the epidemic, prediction of peak timing, and prediction of peak intensity (7). The appraisal concluded that a longer prospective evaluation was needed to ascertain the validity of the results and that data from larger urban counties were required to draw reliable conclusions about generalizability.

In this article, we describe a prospective 10-year evaluation of this local influenza nowcasting method in 3 urban counties (population 1.3–2.2 million) in Sweden. The evaluation period included 1 pandemic (2009) and 9 seasonal influenza epidemics.

Methods

Study Design
We used an open cohort design based on the total population in 3 urban counties: Stockholm County (population 2,231,000), West Gothia County (population 1,649,000), and Scania County (population 1,304,000) (Figure 1). We used retrospective data from January 1, 2008, through June 30, 2009, and prospective data from July 1 through February 28, 2019, from 2 sources in the countywide health information system: daily numbers of clinically diagnosed influenza cases (Figure 2) and daily syndromic chief complaint data from a telenursing service (Figure 3) (6,7). The clinical influenza case data were used to detect the local start of the epidemic and prediction of its peak intensity, and the syndromic data were used to predict the peak timing. Existing evidence of a strong association between the clinical influenza case data and syndromic chief complaint data from the telenursing service was used in this nowcasting method (8,9). Because of a change of system, no syndromic chief complaint data were available for Stockholm County. Syndromic data from West Gothia County were therefore used to predict the peak timing for Stockholm County.

Timeliness was used as a performance metric for detection of the local start of the epidemic and the peak-timing prediction; the correct identification of intensity category on a 5-grade scale (i.e., nonepидemic, low, medium, high, and very high) was used for peak-intensity prediction. The study design was approved by the Regional Research Ethics Board in Linköping (approval no. 2012/104-31).
Definitions
Influenza cases were identified by using codes from the International Classification of Diseases, 10th Revision, for influenza (J10.0, J10.1, J10.8, J11.0, J11.1, J11.8) (10) as recorded in the local electronic health data repository. Only influenza diagnoses in the first coding position were used. Influenza-related telenursing call cases were identified by using the chief complaint codes associated with influenza symptoms. The symptoms used were fever, cough, and headache. These data were downloaded from the electronic patient record systems to the electronic health data repository twice daily.

The intensity level defining the start of a local epidemic (i.e., the intensity that determines the endpoint for the detection function) was set to 6.3 influenza-diagnosis cases/100,000 population recorded during a floating 7-day period in the countywide health information system (6). This level was chosen by inspecting the epidemic curves of previous local influenza epidemics. A recent comparison of influenza intensity levels in Europe estimated a similar level (6.4 influenza-diagnosis cases/week/100,000 population) for the 2008–09 seasonal influenza in Sweden (11). The optimal alerting threshold before each epidemic was decided by calculating the sensitivity and the specificity for the previous nonpandemic influenza seasons and studying them on a receiver operating characteristic curve (6). The calculation of the specificity was based on all days in the nonepidemic period (i.e., before the limit of 6.3 influenza-diagnosis cases/100,000 during a floating 7-day period occur), and the calculation of the sensitivity was based on the days in the epidemic period (i.e., from when the limit of 6.3 influenza-diagnosis cases/100,000 during a floating 7-day period has occurred). Peak timing was defined as the day when the highest number of influenza-diagnosis cases were documented in the countywide electronic patient record. Peak intensity was defined as the number of influenza-diagnosis cases that had been documented at peak timing.

Method Application
Technical details concerning the 3 functions of nowcasting have been described previously (6; Appendix, https://wwwnc.cdc.gov/EID/article/26/11/20-0448-App1.pdf). The functions are detecting the start of the influenza season or pandemic and forecasting the peak day and peak intensity. Once the epidemic has been detected using the clinical influenza data, the syndromic telenursing data are used to detect when it decreases, that being the indication for the peak. Because changes in clinical influenza data have been found to occur 14 days after corresponding changes in syndromic data, the peak timing in the clinical influenza data are forecasted to occur 14 days after the peak in the syndromic data. Finally, the peak intensity is forecasted by using the clinical influenza data. Syndromic data have a higher amplitude, and the relationship between syndromic data and clinical influenza data are not necessarily constant between seasons. Therefore, the clinical data were used to

![Figure 2. Daily numbers of influenza-diagnosis cases per 100,000 population, January 1, 2008–February 28, 2019, in Stockholm County (upper graph), West Gotia County (middle graph), and Scania County (lower graph), Sweden.](image-url)
predict the intensity once the peak day is predicted with the help of syndromic data.

To calibrate the detection component of the nowcasting method, we retrospectively determined weekday effects on recording of influenza-diagnosis cases and a baseline alert threshold by using the retrospective data. These data were collected from January 1, 2008, through June 30, 2009, including 2 non-pandemic influenza seasons (2007–08 and 2008–09). To determine weekday effects, data from the entire retrospective data collection period were used. To determine the initial alert threshold, only data from the seasonal influenza in 2008–09 were used. The 2007–08 seasonal influenza could not be used for this purpose because the season had started before January 1, 2008. Throughout the study period, the calibration data were updated after every seasonal influenza (i.e., no updates of the threshold after the 2009 pandemic outbreak). The detection algorithm was thus applied to the next epidemic by using the revised threshold determined in the updated retrospective dataset.

Before the 2010–11 seasonal influenza, no updates were performed because the set of retrospective data remained the same (i.e., it contained data from the 2008–09 seasonal influenza but excluded pandemic data). For the 2011–12 seasonal influenza, the threshold was updated by using retrospective data from the 2008–09 and 2010–11 seasonal influenza. For the 2012–13 seasonal influenza, the threshold was updated by using retrospective data from the 2008–09, 2010–11, and 2011–12 seasonal influenza, and so on. The weekday effects were assumed to be relatively constant over time in the local detection analyses and were therefore not updated after every seasonal influenza.

The set of retrospective data from the seasonal influenza in 2007–08 and in 2008–09 were also used to initially calibrate peak-timing prediction for West Gothia County and Scania County. The dataset was used to decide the grouping of chief complaints with the largest correlation strength and longest lead time from telenursing data to influenza-diagnosis data (10,11). For both counties, the best performing telenursing chief complaint was fever (among children and adults), and the most favorable lead time was 14 days. When the peak timing had been determined, the second component of the local prediction module was applied to influenza-diagnosis data from the corresponding epidemics to find the peak intensity on the predicted peak day (6). Regarding weekday effects on local prediction, the same calculation was applied and the same grouping of chief complaints and lead time were used throughout the study.

**Metrics and Interpretations**

On the basis of the utility of the nowcasting method in local healthcare settings, the maximum tolerable timeliness error for detection and peak-timing predictions was set to 11 days (≈1.5 weeks). Method performance was defined to be excellent if the absolute value of the timeliness error was <3 days, good if it was 4–7 days, tolerable if it was 8–11 days, and poor if it was >12 days. For the interpretation of peak intensity predictions, the intensity level categories

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**Figure 3.** Daily numbers of telenursing calls attributable to fever (among children and adults) per 100,000 population, January 1, 2008–February 28, 2019, in West Gothia County (upper graph) and Scania County (lower graph), Sweden.
(nonepidemic <0.9, low 0.9, medium 2.4, high 5.5, and very high intensity level 7.9 cases/day/100,000 population) identified using the moving epidemic method for the reference influenza season 2008–09 in Sweden (11) were used. If the predicted peak intensity fell into the same category as the recorded peak intensity, the prediction was considered excellent. If the predicted peak intensity did not fall into the same intensity category, the predicted peak was considered good if it was up to 10% above or below the threshold for the recorded peak intensity category, tolerable if the predicted peak was 10%–20% above or below the threshold for the recorded peak intensity category, and poor otherwise. When assessing series of nowcasts, the performance of a sequence of nowcasts was considered satisfactory if all separate forecasts were assessed as excellent, good, or tolerable, and poor otherwise.

Results

Local Detection
The date of the actual start of the epidemic phase for the 10 influenza epidemics differed by 2–27 days between the 3 counties (Table 1). The detection component of the local nowcasting method showed good or excellent performance in all counties under surveillance for 6 of the 9 nonpandemic influenza seasons and in 2 out of 3 counties under surveillance for the 3 remaining seasons. Twice the poor alerts were issued too soon and once belatedly. The detection performance was good during the 2009 influenza A(H1N1) pandemic in 2 of 3 counties (Stockholm and West Gothia) and poor in 1 county (Scania).

Local Prediction
For the 2009 influenza pandemic, the performance of the peak-timing prediction was poor in all 3 study counties (Table 2). The peak-timing prediction was also poor for the 2010–11 seasonal, when influenza A(H1N1) and B viruses were circulating. Thereafter, the predictions were tolerable for the 2011–12 seasonal influenza, when influenza A(H3N2) virus was circulating, and good to excellent for the remaining influenza seasons, with the exception of the poor peak-timing predictions for Scania County for the 2016–17 and 2017–18 influenza seasons, with influenza A(H3N2) virus circulating in 2016–17 and influenza A(H3N2) and B in 2017–18.

The prediction of the peak-intensity level was poor for the 2009 influenza pandemic in all 3 study counties (Table 2). For seasonal influenza, in 2 of the study counties (Stockholm and West Gothia) the predictions were tolerable to excellent for all seasons, except for the 2018–19 season with influenza A(H1N1) in Stockholm. In 1 county (Scania), the peak-intensity predictions were poor for 5 of the 9 influenza seasons: 2010–11 with influenza A(H1N1) and B, 2011–12 with influenza A(H3N2), 2014–15 with influenza A(H3N2) and B, 2015–16 with influenza A(H1N1) and B, and 2017–18 with influenza A(H3N2) and B circulating.

Discussion
Epidemic forecasts for large administrative areas (e.g., nations or states) might not be sufficiently informative for local response to epidemics if sizable variations in disease transmission patterns exist between the smaller administrative areas (e.g., counties) with independent local healthcare governance that they contain (12). The importance of taking the local context into regard in epidemic forecasting has been further emphasized during the current coronavirus pandemic (13). In our prospective 10-year evaluation of local nowcasting in 3 urban counties, the start of the influenza seasons included differed by up to 27 days and the peak intensity by >1 intensity level among the counties, whereas the time-of-peak differences were small. The purpose of the evaluated local detection function was to allow hospitals and primary healthcare centers time to prepare for management of influenza patients (e.g., by preparing intensive care unit resources or postponing some elective procedures). This component showed satisfactory performance in all 3 counties. The peak-timing prediction function was aimed at informing the local authorities when the peak has occurred and that health service routines soon can be permitted to return to normal. This component showed satisfactory performance from the 2011–12 influenza season onward. Predictions of peak timing were made 8–10 days before the peak and were ±7 days accurate in most cases. This finding contrasts with the current practices in the study counties, where the peak of an influenza season is retrospectively determined from surveillance data ±10–14 days after it has occurred. The nowcasting of peak-intensity level was aimed at warning the local authorities about high-intensity influenza transmission and the potential need for social distancing measures (e.g., closure of kindergartens). This component provided satisfactory information for influenza seasons in 2 out of 3 study counties (Stockholm and West Gothia).

Although the evaluated nowcasting method is automated to run on routinely collected healthcare data, the accuracy of the nowcasts depends on the
stability of the data supply and information infrastructure over time. The method does not require influenza cases to be confirmed by a laboratory as long as data recording remains relatively stable. Nonetheless, some observations can be made about the sensitivity of the local nowcasts to contextual factors. In Sweden, vaccination adapted to the current circulating strains is made available free-of-cost to the elderly and risk groups before every influenza season. However, in the case of the 2009 influenza A(H1N1) pandemic, a national vaccination campaign was implemented, covering the entire population. This intervention probably influenced the nowcasting performance during the corresponding period. Looking only at the performance for seasonal influenza, we observed outcomes in 1 of the 3 study counties (Scania) that raise concerns about vulnerability of the nowcasts to sociodemographic dynamics (14). Malmö (population 450,000; capital of

Table 1. Performance of the detection algorithm displayed with alert thresholds updated by using data from previous nonpandemic influenza seasons in evaluation of nowcasting for detection and prediction of local influenza epidemics, Sweden, 2008–2019

| Influenza virus activity | Updated* alert threshold, cases/day/100,000 population† | Timeliness‡ | Start according to method | Actual start§ | Interpretation |
|--------------------------|--------------------------------------------------------|-------------|---------------------------|---------------|---------------|
| 2009 A(H1N1)             | 0.63                                                   | –5          | 2009 Aug 24               | 2009 Aug 19   | Good          |
|                          | 0.73                                                   | –6          | 2009 Sep 3                | 2009 Aug 28   | Good          |
|                          | 0.25                                                   | 18          | 2009 Aug 13               | 2009 Aug 31   | Poor          |
| 2010–11 A(H1N1) and B¶   | 0.63                                                   | –7          | 2010 Dec 30               | 2010 Dec 23   | Good          |
|                          | 0.73                                                   | –12         | 2011 Jan 9                | 2010 Dec 28   | Poor          |
|                          | 0.25                                                   | 2           | 2010 Dec 23               | 2010 Dec 25   | Excellent     |
| 2011–12 A(H3N2)          | 0.59                                                   | 2           | 2012 Jan 22               | 2012 Jan 24   | Excellent     |
|                          | 0.43                                                   | 1           | 2012 Jan 31               | 2012 Feb 1    | Excellent     |
|                          | 0.27                                                   | 23          | 2012 Jan 9                | 2012 Feb 1    | Poor          |
| 2012–13 A(H3N2), A(H1N1), and B | 0.51                                      | –6          | 2013 Jan 3               | 2012 Dec 28   | Good          |
|                          | 0.44                                                   | 0           | 2012 Dec 29               | 2012 Dec 29   | Excellent     |
|                          | 0.28                                                   | 0           | 2012 Dec 27               | 2012 Dec 27   | Excellent     |
| 2013–14 A(H3N2), A(H1N1), and B | 0.52                                      | 0           | 2014 Jan 30               | 2014 Jan 30   | Excellent     |
|                          | 0.37                                                   | 1           | 2014 Jan 27               | 2014 Jan 27   | Excellent     |
|                          | 0.35                                                   | 0           | 2014 Jan 28               | 2014 Jan 28   | Excellent     |
| 2014–15 A(H3N2) and B     | 0.52                                                   | –6          | 2015 Jan 13               | 2015 Jan 7    | Good          |
|                          | 0.39                                                   | 0           | 2015 Jan 17               | 2015 Jan 17   | Excellent     |
|                          | 0.35                                                   | 7           | 2015 Jan 16               | 2015 Jan 23   | Good          |
| 2015–16 A(pH1N1) and B    | 0.52                                                   | 0           | 2016 Jan 2                | 2016 Jan 2    | Excellent     |
|                          | 0.47                                                   | 16          | 2015 Dec 28               | 2016 Jan 13   | Poor          |
|                          | 0.34                                                   | 0           | 2015 Dec 16               | 2015 Dec 16   | Excellent     |
| 2016–17 A(H3N2)          | 0.34                                                   | –2          | 2016 Dec 1                | 2016 Nov 29   | Excellent     |
|                          | 0.31                                                   | –2          | 2016 Dec 17               | 2016 Dec 15   | Excellent     |
|                          | 0.31                                                   | 0           | 2016 Dec 10               | 2016 Dec 10   | Excellent     |
| 2017–18 A(H3N2) and B     | 0.38                                                   | 0           | 2017 Dec 12               | 2017 Dec 12   | Excellent     |
|                          | 0.44                                                   | 4           | 2017 Dec 30               | 2018 Jan 3    | Good          |
|                          | 0.34                                                   | 5           | 2017 Dec 22               | 2017 Dec 27   | Good          |
| 2018–19 A(pH1N1)         | 0.36                                                   | –7          | 2018 Dec 18               | 2018 Dec 5    | Good          |
|                          | 0.40                                                   | –6          | 2018 Dec 28               | 2018 Dec 22   | Good          |
|                          | 0.34                                                   | 5           | 2018 Dec 27               | 2019 Jan 1    | Good          |

*Threshold updated after every seasonal influenza (i.e., no updates after pandemic outbreaks).
†Threshold determined using clinical influenza-diagnosis data.
‡Positive value means that the algorithm issued an alarm before the local epidemic had started; negative value means that the alarm was raised after the start of the epidemic.
§Actual start is the date when the retrospectively calculated intensity level reached the predefined threshold for start of an epidemic (6.3 influenza-diagnosis cases/100,000 population recorded during a floating 7-day period) (7,11).
¶No update of threshold before this seasonal influenza because the previous outbreak was a pandemic.
Scania County, Sweden) and Copenhagen (population 2 million; capital of Denmark) are connected by a bridge providing for daily commuting between the metropolitan areas, and their labor markets are closely integrated. The epidemic situation in the highly cosmopolitan Copenhagen region might have had a stronger influence on influenza epidemics in Scania County than the epidemic situation in the neighboring regions had on the other study counties. By structured introduction, evaluation, and modification of prediction models that use additional data sources and statistical methods, local nowcasting can be adapted also to communities with unusual characteristics (15,16). This evidence-based strategy means that our method can be incrementally adapted to modeling of, for instance, local rural or semirural communities in which residents commute extensively to a neighboring city that is not included in the model.

Some possible limitations exist in terms of the design of this prospective evaluation that require attention. First and foremost, whether the framework used to interpret the nowcasting performance is adequate from the local health authority perspective should be assessed. Regarding the time-of-peak

| Influenza virus activity | Prediction date | Predicted | Error | Interpretation | Predicted | Factual | Interpretation |
|-------------------------|----------------|-----------|-------|----------------|-----------|---------|----------------|
| 2009 A(H1N1)            |                |           |       |                |           |         |                |
| Stockholm               | 2009 Sep 13    | 8         | 56    | Poor           | Medium (5.0) | Very high (12.4) | Poor |
| West Gotha              | 2009 Sep 13    | 8         | 56    | Poor           | Low (2.2)   | Very high (13.7) | Poor |
| Scania                  | 2009 Sep 25    | 10        | 42    | Poor           | Low (1.4)   | High (6.4)     | Poor |

| 2010–11 A(H1N1) and B   |                |           |       |                |           |         |                |
| Scania                  | 2011 Jan 14    | 10        | 28    | Poor           | Medium (3.4) | Medium (3.5) | Excellent |
| West Gotha              | 2011 Jan 14    | 10        | 14    | Poor           | Medium (4.3) | High (6.1)  | Tolerable |
| Scania                  | 2011 Jan 10    | 11        | 22    | Poor           | Medium (2.9) | High (5.5)  | Poor     |

| 2011–12 A(H3N2)         |                |           |       |                |           |         |                |
| Stockholm               | 2012 Feb 27    | 8         | –8    | Tolerable      | High (7.4) | Very high (9.4) | Good |
| West Gotha              | 2012 Feb 27    | 8         | –8    | Tolerable      | High (7.8) | Very high (9.6) | Good |
| Scania                  | 2012 Feb 27    | 8         | –8    | Tolerable      | Medium (4.0) | High (6.8)  | Poor    |

| 2012–13 A(H3N2), A(H1N1), and B |                |           |       |                |           |         |                |
| Stockholm               | 2013 Feb 10    | 8         | –7    | Good           | Very high (10.3) | Very high (12.2) | Excellent |
| West Gotha              | 2013 Feb 10    | 8         | –7    | Good           | Very high (10.3) | Very high (11.9) | Excellent |
| Scania                  | 2013 Feb 8     | 10        | –7    | Good           | High (7.3)   | Very high (10.7) | Good    |

| 2013–14 A(H3N2), A(H1N1), and B |                |           |       |                |           |         |                |
| Stockholm               | 2014 Feb 16    | 8         | –7    | Good           | Medium (2.7) | Medium (3.0) | Excellent |
| West Gotha              | 2014 Feb 16    | 8         | –7    | Good           | Medium (3.5) | Medium (2.9) | Excellent |
| Scania                  | 2014 Feb 17    | 8         | –1    | Excellent      | Medium (3.2) | Medium (4.2) | Excellent |

| 2014–15 A(H3N2) and B   |                |           |       |                |           |         |                |
| Stockholm               | 2015 Feb 22    | 8         | 6     | Good           | Medium (4.5) | High (6.5)  | Tolerable |
| West Gotha              | 2015 Feb 22    | 8         | 6     | Good           | Very high (7.9) | Very high (8.3) | Excellent |
| Scania                  | 2015 Feb 14    | 9         | 0     | Excellent      | Medium (3.9) | Very high (8.1) | Poor    |

| 2015–16 A(H1N1) and B   |                |           |       |                |           |         |                |
| Stockholm               | 2016 Feb 7     | 8         | 0     | Excellent      | High (6.7)  | Very high (8.2) | Tolerable |
| West Gotha              | 2016 Feb 7     | 8         | 7     | Good           | High (7.6)  | Very high (11.6) | Good    |
| Scania                  | 2016 Feb 6     | 9         | 7     | Good           | Medium (4.3) | Very high (10.4) | Poor    |

| 2016–17 A(H3N2)         |                |           |       |                |           |         |                |
| Stockholm               | 2017 Jan 1     | 8         | –7    | Good           | Very high (8.2) | High (6.8)  | Good    |
| West Gotha              | 2017 Feb 12    | 8         | 7     | Good           | Medium (3.3) | Medium (3.7) | Excellent |
| Scania                  | 2017 Feb 5     | 8         | 14    | Poor           | Medium (4.2) | Medium (5.1) | Excellent |

| 2017–18 A(H3N2) and B   |                |           |       |                |           |         |                |
| Stockholm               | 2018 Feb 18    | 8         | –7    | Good           | Very high (14.4) | Very high (11.6) | Excellent |
| West Gotha              | 2018 Feb 18    | 8         | 0     | Excellent      | Medium (5.2) | High (5.9)  | Good    |
| Scania                  | 2018 Feb 4     | 8         | 14    | Poor           | Medium (4.2) | Very high (14.0) | Poor    |

| 2018–19 A(H1N1)         |                |           |       |                |           |         |                |
| Stockholm               | 2019 Feb 3     | 8         | 0     | Excellent      | Very high (14.4) | High (6.2)  | Poor    |
| West Gotha              | 2019 Feb 3     | 8         | 7     | Good           | Medium (4.0) | Medium (3.4) | Excellent |
| Scania                  | 2019 Feb 3     | 8         | –7    | Good           | Medium (2.8) | Medium (5.2) | Excellent |

† Time-to-peak (days) determined using syndromic telenursing data. Positive value means that the peak was predicted to be reached before the actual peak occurs, whereas negative value means that the peak is predicted after the actual peak occurs.

‡ Peak-intensity category determined using clinical influenza-diagnosis data.

§ Using clinical influenza data (Table 1; https://wwwnc.cdc.gov/EID/article/26/11/20-0448-T1.htm), the start of the epidemic was detected on December 27. On February 1, using syndromic data, the peak in clinical influenza data was forecasted to occur 8 days later (February 9), but the peak actually occurred on February 2 (7 days earlier than forecasted). Also, on February 1, the clinical influenza data intensity was forecasted to be high.
predictions, the ongoing FluSight study uses weekly data (4), thus accepting forecasts made at a weekly resolution. The evaluation framework used to classify forecasts as excellent was at a higher temporal resolution (less than one half week). This boundary was defined from a county government perspective, where the attention is on local resource allocation (e.g., intensive-care unit facilities and hospital beds) for the care of influenza patients. In this situation, nowcasts that are off by days to weeks might have severe consequences for patients in need of these resources. Categories that are suitable for evaluation of usefulness in local response preparations might not be suitable for interpretation of utility in national or international response planning. These observations suggest that the requirements on the accuracy of peak-timing predictions are context-dependent and warrant further research. Concerning the predictions of peak intensity, evaluation of the peak-intensity forecasts indicated that 22% (6 of 27) of the seasonal influenza nowcasts were poor. Retrospectively documenting baseline and threshold values for influenza epidemics helps define whether an influenza epidemic has been different in intensity compared with previous seasons and thereby contributes to future preparedness planning (17,18). For the evaluation of intensity predictions in this study, we used the thresholds established using the modeling epidemic method from the reference 2008-09 seasonal influenza season. To improve the validity of the assessments, annual updates of the threshold values using county-level data from previous seasons should be considered for future evaluations of local influenza nowcasting.

Longitudinal prospective evaluations might be needed to draw valid conclusions concerning the performance of local epidemic nowcasting, and inclusion of data from urban counties might be required for generalizability (7). We found in our study that the performance of seasonal influenza nowcasting was satisfactory during a 10-year period in 3 urban counties but poorer in 1 county, possibly because of sudden sociodemographic changes. We conclude that the performance of the local nowcasting method was satisfactory for seasonal influenza. The results are of general interest for local healthcare planning during epidemics because the precision by which healthcare systems can adapt its resources to the management of infected patients in these situations affects the resource availability for all other patient groups.

This study was supported by grants from the Swedish Civil Contingencies Agency (grant no. 2010-2788) and the Swedish Research Council (grant no. 2008-5252). The funders had no role in the study design, data collection and analysis, decision to publish, or preparation of the manuscript.

Authors’ contributions: A.S., O.E., Ö.D., B.J.C., M.B., and T.T. conceived and designed the study; A.S., O.E., and Ö.D. analyzed the data; A.S., O.E., Ö.D., G.L., and T.T. contributed materials and analysis tools; A.S. and T.T. wrote the paper; Ö.D., O.E., B.J.C., M.B., G.L., A.J., and E.I. revised the manuscript and provided intellectual content; and A.S., O.E., Ö.D., B.J.C., M.B., G.L., A.J., E.I., and T.T. gave final approval of the version to be published. T.T. is guarantor of the content.

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Nowcasting (Short-Term Forecasting) of Influenza Epidemics in Local Settings, Sweden, 2008–2019

Appendix

Method Design Overview

The nowcasting method is divided into separate modules for detection and prediction of influenza activity, respectively (1). An overview of the main statistical assumptions and equations for each method component is displayed in Appendix Figure. The function of the detection module is to alert for a period of increased load of influenza-diagnosis cases on local health care services, whereas the function of the prediction module is to predict the timing of the activity peak and its intensity. The prediction process is divided into two components. In the first component, syndromic data are used to predict the peak timing, and in the second component, influenza-diagnosis data are used to estimate the peak intensity.

In the study setting, patients clinically diagnosed with influenza were used as the endpoint measure. Early detection of increased influenza activity and prediction of peak intensity were thus based on streams of the endpoint measure data, whereas prediction of peak timing was based on syndromic data.

Detection Module

Exponential regression (1) is used for detection modeling, based on the observation that the beginning of an influenza epidemic is assumed to have an exponential growth of infected individuals;

\[ X_t = e^{a_0 + b_1 t}, \]

with \( t \) representing the time, \( a_0 \) representing the level, and \( b_1 \) representing the trend. The expected number of visits at local health care services, \( E[Y_t] \) is the product of \( X \) and the
probability $p$ for an infected individual to visit the local health care service. This expectation is also exponential in time;

(2) $E[Y_t] = e^{a_0 + b_1 t} p = e^{a_0 + \ln(p) + b_1 t} = e^{b_0 + b_1 t},$

Where $b_0$ now combines the current level of number of infected and probability of visiting the local health care service without any possibility to separate them. As daily data are used in the analysis, weekday effects, $A_w$, are also calculated and used as an offset variable in the exponential regression analysis. The weekday effects are calculated as follows: let $A_{Monday}$ be the average number of events on Mondays during previous epidemics and denote the values for other weekdays by $A_{Tuesday}$, $A_{Wednesday}$, and so on. Let $A_{Total} = (A_{Monday} + ... + A_{Sunday})/7$. The multiplicative weekday effect for Mondays is $A_{Monday}/A_{Total}$ and so on. The weekday effects are included in the model;

(3) $E[Y_t] = e^{b_0 + b_1 t + \ln(A_w)},$

If $X$ is large, $p$ is small, and the infected individuals act independently, then $Y$ is approximately Poisson distributed;

(4) $Y_t \sim Poisson (e^{b_0 + b_1 t + \ln(A_w)}).$

Furthermore, the time is shifted, that is, the most recent day is considered as $t = 0$, the second most recent day is considered as $t = -1$, and so on. For every new day, the time axis is moved one step so that the new “most recent day” is considered as $t = 0$. For each day an exponential regression analysis (1) is run and a fitted value $\hat{y}$ is calculated by inserting $t = 0$ in equation (3) giving

(5) $Y_t = e^{b_0},$

as an estimate of the current level of visits which is smoothed for random variation and adjusted for weekday effects. This is repeated for each day by moving the time axis 1 day at a time so that the most recent point in time of the series is considered $t = 0$. Doing this, one value is obtained for every day representing the level for that day. Finally, the lower 95% confidence limit is calculated to represent the level of influenza activity, which is then compared with a predetermined threshold. If the level (i.e., the lower 95% confidence limit) is above the threshold, an alarm is raised, which means that the winter influenza season (or pandemic) has started; and if the level is below the threshold, no alarm is raised.
Detection starts when the previous epidemic has ended (the interepidemic period level for the community where the detection component is applied), and runs during the inter-epidemic period until an increase in diagnosed influenza cases is detected. When the increase is confirmed, the algorithm is paused and restarted when the epidemic has ended.

**Prediction Module**

**Peak-Timing Prediction**

In the first component, the aim is to predict the peak timing using linear regression. Including weekday effects $A_w$ and smoothed for random variation, the model for the number of cases in syndromic data are expressed as

\[ Z_t = (b_0 + b_1 t) \times A_w, \]

with $b_0$ representing the level and $b_1$ representing the trend. Since the weekday effects $A_w$ are known, a model smoothed for weekday effects and random variation can be expressed as

\[ Z_t / A_w = b_0 + b_1 t \]

For each 7-day period, a linear regression (7) is run and parameter estimates $b_0$ and $b_1$ are fitted. The idea is to estimate the trend in syndromic data for every 7-day period (the first period being days 1–7 and the second being days 2–8), from the beginning of an epidemic and until the peak is found. Although it is unlikely that an epidemic curve increases and decreases linearly, the assumption can be made that the trend during a short period of 7 days has almost a linear increase or decrease.

The search for the peak starts when the detection algorithm signals that an epidemic has taken off and continues until the peak is detected. To identify the peak timing, two conditions are set. As per the first condition, it is essential to ensure that the epidemic has a sufficiently sharp upward trend. The trend is therefore defined as sufficiently sharp when significantly positive ($p < 0.10$) trends $b_1$ have occurred either during two consecutive or during three different 7-day periods. When one of these events has occurred, the second condition is applied. According to this condition, when either 1) the first significantly negative trend ($b_1$) during a 7-day period has occurred or when 2) negative trends ($b_1$) (regardless if they are significant or not) have occurred during five of the latest seven 7-day periods, it is assumed that the peak has been reached on the
first day of this period. However, there is a possibility that these 7-day periods “overlaps” with a previous 7-day period, which includes a significantly positive trend. In that case, the first 7-day period with a significantly negative trend is ignored and the peak is instead assumed to appear during a forthcoming period which does not include periods with significantly positive trends. The search is aborted if the peak is not found when the epidemic has already descended in the local setting where the algorithm is applied.

When the peak is found in the syndromic data, the clinical influenza case data for the succeeding 14 days (2) are used to find the peak in clinically diagnosed influenza cases. In other words, if the peak in the syndromic data appears on day 0, the influenza-diagnosis peak is assumed to appear on day 14. However, it is possible that the peak in the syndromic data occurs on a day during the weekend, while it is unlikely that the peak in influenza-diagnosis data occurs on one of these days due to lower healthcare access, e.g., primary care centers are closed during weekends in Sweden. Instead, it is reasonable to assume that the influenza-diagnosis peak occurs at the beginning of the week because individuals who suffer influenza symptoms during the weekend visit primary care centers when they reopen on Monday or possibly Tuesday. Adjustments are therefore made by moving the influenza-diagnosis peak to the following Monday if it is expected to occur on a Friday, Saturday, or Sunday according to syndromic data and to the previous Tuesday if the peak is expected to take place on a Wednesday or Thursday. If the peak is expected to occur on a Monday or Tuesday, no adjustments are made. In other words, in the first case the syndromic data precedes influenza-diagnosis data between 15 and 17 days, in the second case between 12 and 13 days, and in the third case 14 days.

Depending on what day of the week the peak in the syndromic data are expected to take place, the prediction of the influenza-diagnosis peak is made between 6 and 11 days before it is expected to occur, as the syndromic peak can be determined first after 6 days has passed of the syndromic data series.

**Peak-Intensity Prediction**

In the second component of the prediction module, the aim is to predict only the peak intensity. Based on empirical assessments of previous epidemics, an epidemic adjusted for weekday effects is assumed to show a bell-shaped form from the beginning to the end, and can therefore be expressed using a derivate of a normal distribution density function. The intensity
function must also include weekday effects and total number of events during the whole epidemic. Assuming that the peak timing is known (estimated in the first prediction component) and that an epidemic follows the bell-shaped function around the peak, the intensity function can be used to predict the peak intensity at time $m$.

Assume that day number $t = 1, 2, 3, ..., t_i$; the observed number of influenza-diagnosis cases is $y = y_1, y_2, y_3, ..., y_i$, and that

\begin{align}
(8) \quad Y_t \sim \text{Poisson} \left( T \times w \times f(t; m, s) \right),
\end{align}

where $T$ is the total number of health care visits of the whole epidemic, $w$ is the weekday effects, $f$ is the normal distribution density function, $t$ is the day number, $m$ is the center of the epidemic (which coincides with $t$ for the peak), and $s$ is the spread in time. Since $t$, $w$, and $m$ are known, only the parameters $T$ and $s$ are estimated using $y$ in such way so that the likelihood is maximized. However, to do that, first appropriate starting values for these parameters need to be selected. Finally, using the known parameter $m$ and the estimated parameters $T$ and $s$, the peak intensity at time $m$ is calculated by replacing $t$ with $m$ in equation (8).

It is important that the start of the series seems appropriate because the second prediction component assumes that the level is zero or at an interepidemic level at the start and it is not optimal that there are single or occasional spikes at the beginning of the series. For that reason, the start of the series should be a couple of weeks before an epidemic is detected.

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Appendix Figure. Overview of the main mathematical equations or functions used for each component.