Psychometric Properties and Factor Structure of the Brief Religious Coping Scale (Brief-RCOPE) in Puerto Rican Adults

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Abstract
This research examines the psychometric properties and factor structure of the Brief Religious Coping Scale (Brief-RCOPE) in a sample of 302 Puerto Rican adults. We examined its internal consistency (Cronbach’s alpha), McDonald’s omega coefficient, construct validity, and factor structure. The Brief-RCOPE obtained an alpha coefficient of .94 in the Positive Religious Coping (PRC) subscale and .84 in the Negative Religious Coping (NRC) subscale. The omega coefficient was .94 (PRC) and .85 (NRC), respectively. We conducted a confirmatory factor analysis, using the Satorra-Bentler correction, to examine the factor structure of the Brief-RCOPE. The two-factor model showed a better adjustment to the data than the one-factor model. Indicators of construct validity were also adequate. Our findings suggest that the Brief-RCOPE is a reliable and valid instrument to measure religious coping strategies that may significantly affect people’s daily lives.

Keywords: factor structure, Puerto Rican, psychometric properties, religious coping, religiousness
Introduction

Research studies reveal that religious/spiritual involvement is associated with better health (Koenig, 2012, 2015; Oman & Syme, 2018). The documented findings concerning the impact of religious/spiritual involvement on health are closely associated with the use of religious coping (RC) strategies (Gerber, Boals, & Schuettler, 2011; Gonçalves, Lucchetti, Menezes, & Vallada, 2015; Pargament, 1997). RC is a strategy based on religious beliefs and practices to prevent and alleviate the negative consequences of stressful events (Pargament, 1997).

Religiosity is a relevant dimension of Puerto Ricans’ culture and lifestyle. According to the Pew Research Center survey (2014), 89% of Puerto Ricans living on the island self-perceive as Christians, distributed in Catholic Christians with 56%, followed by Protestant Christians with 33%, while 8% are unaffiliated and 2% who identify themselves as “other”, which could include minority religious groups. Therefore, it is not surprising that religious and spiritual beliefs play a significant role in the majority of Puerto Ricans’ daily life and culture (Agosto-Cintrón, 1996; Scarano, 2008).

There are several reasons to argue the importance of measuring the religious and spiritual dimensions. First, research reveals that these dimensions can affect people’s health in physical, emotional and social aspects when used as positive or negative coping mechanisms (Bonelli & Koenig, 2013; Koenig, 2012; Oman & Syme; 2018). Second, on many occasions, the main complaint of patients/clients who attend therapy is related to religious/spiritual aspects (American Psychiatric Association, 2013). Therefore, the measurement of these constructs will provide more information on the role of religious and spiritual dimensions in people’s lives. Third, a close examination of religiosity and religious coping strategies could help us understand the global vision of the patient/client and strengthen the therapeutic relationship (Richards & Bergin, 2014). Fourth, the information obtained could be relevant in the development of a treatment plan consistent with the needs of the client/patient (Richard & Bergin, 2014). Last, the assessment of these dimensions could help professionals understand the role of spirituality and religiosity in client/patient health care (Gonçalves et al., 2015).

Pargament (1997) defines coping as “the search for meaning in times of stress” (p. 90). In addition, religious coping is defined as the different ways of understanding and handling negative life events that are related to the sacred (Pargament & Raiya, 2007). Pargament (1997) originally developed the religious coping construct. This author proposes that religion is one of the ways in which individuals can cope with their life situations through positive and negative strategies that emerge from their religious beliefs and practices. Pargament, Koenig and Perez (2000) developed the first validated instrument to measure religious coping: The Religious Cope (RCOPE). This scale, in its original form, had 105 items distributed in 21 sub-scales. As reported in the study, the reliability estimates of the instrument subscales were high. Specifically, the RCOPE showed a Cronbach alpha internal consistency of .80 or more for all subscales except for two dimensions: marking the religious limits (.78) and the reassessment of the power of God (.61)

Later, a short version of the RCOPE was developed. The Brief Religious Coping Scale (Brief-RCOPE) includes 14 items as a result of conducting several exploratory factor analyses (EFA) with different samples (Pargament, Smith, Koenig, & Perez, 1998; Pargament, Feuille, & Burdzy, 2011). This measure has the advantage of measuring religious coping strategies in a short
time. The scale consists of 14 items in four-point Likert-type format ranging from *Not at all* to *Very much*. The items are distributed in two dimensions classified as positive religious coping (PRC) and negative religious coping (NRC), as shown by both EFA and a confirmatory factor analysis (CFA) conducted by the authors on two samples (Pargament et al., 1998). The model fit indexes obtained for the CFA two-factor model were adequate. In their review, Pargament et al. (2011) reported that Cronbach’s alphas for the NRC were generally lower than those for PRC, with median values for the PRC scale being .92 and .081 for the NCR.

The psychometric properties and factor structure of the Brief RCOPE have been examined in a diversity of countries and populations. For instance, in a sample of 403 Iraqi secondary school students, and following a principal component analysis (PCA), with both varimax and oblimin rotations, PRC and NRC subscales had Cronbach’s alphas of .86 and .82, respectively (Al-Hadethe, Hunt, Thomas, & Al-Qaysi, 2016). Also using a PCA but with a promax rotation, Mohammadzadeh and Najafi (2016) examined the structure of the Persian version of the Brief RCOPE among 339 Iranian university students (mean age = 27.30 years). They reported alpha coefficients of .79 for the PRC and .71 for the NRC components extracted, using the eigenvalue > 1.0 criteria. The components showed the same item organization as the original English version. In these two previous studies, the authors did not report the correlation among the components or the observed subscales scores. On the other hand, in three Greek-Orthodox samples, and using an EFA with unweighted least squared extraction and promin oblique rotation, the B-RCOPE showed a two-dimensional factor structure with remarkable stability across the samples corresponding to the PRC (Factor 1) and NRC (Factor 2) dimensions. Cronbach’s alphas were .91 - .96 and .77 - .92 for the PRC and NRC factors, respectively (Paika et al., 2017). The authors reported factor inter-correlations ranging from .33 to .51 in the sub-samples, and a value of .44 for the entire sample. In 2 out of 3 sub-samples, and in the combined sample, item 13 (demonic reappraisal) showed higher loadings on Factor 1, although a higher loading on Factor 2 was expected.

The complexity of the loadings of item 13 was also documented in a study with 170 Brazilian adults with end-stage renal disease, in which the PRC (α = .83) and NRC (α = .75) dimensions were identified after a PCA with varimax rotation (Ramirez et al., 2012). No data was reported on components inter-correlation. In another study conducted in Brazilian adults (Esperandio, Escudero, Fernandes, & Pargament, 2018), the authors split the sample in two: one for conducting an EFA (n = 249) and the other (n = 276) to perform a CFA. In the EFA sample (principal axis factoring extraction with varimax rotation), a two-factor solution was reported, with alpha coefficients of .89 and .85 for the PRC and the NRC, respectively. This factor structure was tested with a CFA in the second sample and adequate model fit indexes were observed. In addition, an average variance extracted of .50 (minimum size recommended) was found for each factor, with a composite reliability of .87 for the PRC and .84 for the NRC. Data on the inter-factor correlation was not provided, nor the path diagram of the CFA, although the factors were considered orthogonal in the EFA. In a third study conducted with a Brazilian Portuguese version, Freitas et al. (2015) used the Brief RCOPE in 147 adults (73.5% Roman Catholics) with inflammatory bowel disease. The authors conducted a PCA with varimax rotation and an eigenvalue > 1.5 as the criteria for component retention. Two components were retained which were consistent
with the PRC (Factor 1, $\alpha = .87$) and NRC (Factor 2, $\alpha = .74$) dimensions. Item 14 (Questioned the power of God) showed the lowest loading with its respective component (.31). The authors kept this item even when it did not meet their cut-off criteria (a loading $\geq .40$).

Spanish versions of the Brief-RCOPE have been used at least for the past 14 years. As far as can be ascertained, the first Spanish version of this measure was developed by Rivera-Ledesma and Montero-López (2007). In two samples (sample 1, $n = 129$; sample 2, $n = 209$) of Mexican adults aged 50 and over (88% Catholics), these authors found internal consistency (alpha) values ranging from .82 to .83 for the PRC and from .60 to .65 for the NRC. When they removed item 13 from the NRC subscale, its alpha values were .62 and .67, respectively. The authors then submitted data from the larger sample to a PCA with varimax rotation, using the eigenvalue $> 1.0$ criteria to determine the number of components. Although they found four components in the initial analysis, only the PRC showed a configuration identical to the original version. Only three items loaded on the NRC component, yielding an alpha coefficient of .50. The internal consistency of the PRC ($\alpha = .83$) and NRC ($\alpha = .61$) was also reported in a study in which Robles-García et al. (2014) used the Spanish Brief RCOPE in a sample of Mexican patients with paranoid schizophrenia. However, in this study, the factor structure of the scale was not examined. Martinez and Sousa (2011), on their behalf, used a Spanish version of the scale in a sample of 121 Mexican-American adults with type 2 diabetes (82% Catholics). In their first PCA, they found three factors with eigenvalues $> 1.0$, with the third factor being composed by items 6 and 7. After excluding those items, additional PCAs (with oblimin and varimax rotations) revealed a two-component solution in which item 13 did not load onto any of them. Cronbach’s alpha values for the 5-tem PRC and the 6-item NRC were of .85 and .86, respectively.

However, the Brief-RCOPE bifactor structure was replicated in a sample of 442 Spanish-speaking Chileans aged 18 to 83 years who had been exposed to traumatic events (García, Oyanedel, Páez, & Arias, 2021). The measure obtained a Cronbach’s alpha of .94 for the PRC subscale and .79 for the NRC dimension. In this study, the authors used a CFA with the robust weighted least square estimation due to the lack of multivariate normality of the data. Model fit indexes yielded excellent results. The correlation among the PRC and NRC factors was .35. In a fifth study conducted with a Spanish version of the scale, Mezzadra and Simkin (2017) used the Brief RCOPE with 200 Catholic students from Buenos Aires, Argentina (aged 14 to 18 years). The authors analyzed data with a polychoric correlation matrix using a CFA and obtained adequate goodness of fit indexes for a two-factor structure. Alpha coefficients for the PRC and NRC were .83 and .72, respectively.

The scientific study of religion and spirituality from a mental health perspective in Puerto Rico has increased in recent years (González-Rivera et al., 2019; Pagán-Torres, Sánchez-Galarza, Tollinchi-Natali, & González-Rivera, 2017). Currently, there are wide varieties of religious and spiritual measures validated with Puerto Rican samples (Pagán-Torres & González-Rivera, 2019). Recently, González-Rivera and Pagán-Torres (2018) validated in Puerto Rico a religious coping scale with 350 adult participants. The measure obtained a Cronbach’s alpha internal consistency of .95. This scale is based on the Lazarus and Folkman (1986) Transactional Model of Stress and Coping, which conceptualizes coping styles in two dimensions: internal and external coping strategies. The first study in which a Spanish version of the Brief RCOPE was used in
Puerto Rico was conducted with a sample of 70 Puerto Rican adult patients (61% Catholics) with cancer (Rodríguez-Carrión, Sayers-Montalvo, & Martínez-Taboas, 2011). Nevertheless, no data about the psychometric performance of the scale in that sample was provided. Years later, Colón-Rivera (2014) translated into Spanish and validated the Brief RCOPE with 226 Puerto Rican adults. The psychometric properties of the instrument revealed a reliability coefficient of .93 for the PRC subscale, and .88 for the NRC subscale. However, the factor structure of the Brief RCOPE has not been explored in a sample of Puerto Ricans.

Therefore, this study has the following aims. First, (a) to examine the factor structure (unidimensional or multidimensional) of the Brief RCOPE, using CFA with the maximum likelihood estimation, in a sample of Puerto Rican adults, given that the state of the research literature reflects a lack of evidence on the factor structure of the Brief RCOPE in Puerto Ricans (Pagán-Torres & González-Rivera, 2019). Second, (b) to examine the reliability of the Brief RCOPE through Cronbach’s alpha internal consistency and McDonald’s omega coefficient. Third, (c) to evaluate the corrected item-total correlation of each item and concurrent validity of each subscale. Finally, (d) to examine the construct validity, through the evidence of its convergent and discriminant validity, using the average variance extracted (AVE) and related statistics.

Method
Research design and procedures

This research has an instrumental design consisting of a one-time assessment. This is a secondary analysis from a research study authorized by the Institutional Review Board (IRB) from Ponce Health Sciences University, Ponce, Puerto Rico (protocol #1902005352). Once the IRB authorization was obtained, the recruitment of the participants began. The digital platform PsychData was employed to collect the online survey data. Online recruitment was achieved through the sharing of study information via social networks and emails. When the participants accessed the survey link, they proceeded to read the informed consent form, which explained all the information, the purpose, the procedures, and the benefits and risks of the research. If the participants agreed to participate, they proceeded to communicate their consent in the space provided in the digital form. In order to guarantee the protection of confidentiality, only an identification code was assigned in the database to record the data of the participants, but no identifying data was collected. After completing the informed consent form, participants proceeded to complete the sociodemographic data form, as well as the study measures.

Participants

A non-probabilistic recruitment strategy was applied. The convenient sample consisted of 302 Puerto Rican adults. The sample average age was 35.79 years (SD = 12.14). The inclusion criteria were: (1) being 21 years of age or older, (2) possessing the ability to read and understand Spanish, (3) being Puerto Rican, and (4) being a resident of Puerto Rico. Table 1 shows the full sociodemographic characteristics of the participants.

Measures

Sociodemographic Data Form. This document included questions aimed to explore the profile of
Table 1
Sociodemographic Characteristics of the Sample.

| Demographics              | f   | %    |
|---------------------------|-----|------|
| **Sex**                   |     |      |
| Male                      | 78  | 25.8 |
| Female                    | 224 | 74.2 |
| **Age**                   |     |      |
| 21-29                     | 134 | 44.2 |
| 30-39                     | 66  | 21.9 |
| 40-49                     | 48  | 15.9 |
| 50-59                     | 41  | 13.8 |
| 60-69                     | 11  | 3.6  |
| 70-71                     | 2   | 0.6  |
| **Marital Status**        |     |      |
| Single                    | 148 | 49.0 |
| Married                   | 102 | 33.8 |
| Widowed                   | 3   | 1.0  |
| Divorced                  | 18  | 6.0  |
| Cohabitng (free union)    | 31  | 10.3 |
| **Annual Income (USD)**   |     |      |
| $0–20,000                 | 150 | 49.7 |
| $21,000–30,000            | 47  | 15.6 |
| $31,000–40,000            | 34  | 11.3 |
| $41,000–50,000            | 13  | 4.3  |
| $51,000–60,000            | 18  | 6.0  |
| $61,000 or more           | 40  | 13.2 |
| **Academic Preparation**  |     |      |
| High school or less       | 14  | 4.6  |
| Associate degree/technical| 13  | 4.3  |
| Bachelor’s degree         | 84  | 27.8 |
| Master’s degree           | 102 | 33.8 |
| Doctoral degree           | 89  | 29.5 |

| Religious Affiliation     |     |      |
| Catholic                  | 124 | 41.1 |
| Protestant (Evangelical,  |     |      |
| Methodist, Baptists,      |     |      |
| Pentecostal)              | 116 | 38.4 |
| Adventist                 | 2   | 0.7  |
| Islamism (Muslim)         | 1   | 0.3  |
| Buddhism                  | 5   | 1.7  |
| Santeria                  | 1   | 0.3  |
| None                      | 53  | 17.5 |

| Importance of Religion    |     |      |
| Nothing                   | 36  | 11.9 |
| Somewhat                  | 57  | 18.9 |
| Important                 | 80  | 26.5 |
| Very important            | 129 | 42.7 |

| Participation in religious activities |     |      |
| Never                                 | 59  | 19.5 |
| Once a year                           | 79  | 26.2 |
| Monthly                               | 44  | 14.6 |
| Weekly                                | 107 | 35.4 |
| Daily                                 | 13  | 4.3  |

| Participation in private religious activities |     |      |
| Never                                 | 60  | 19.9 |
| Once a year                           | 28  | 9.3  |
| Monthly                               | 25  | 8.3  |
| Weekly                                | 53  | 17.5 |
| Daily                                 | 136 | 45.0 |

**Note.** N = 302.
the study participants such as the age, marital status, gender, annual income, religious affiliation, importance assigned to religion, participation in religious activities, and participation in private religious practices.

**Brief Religious Coping Scale (Brief-RCOPE; Pargament et al., 1998).** To measure religious coping, we used the Brief Scale of Religious Strategies (Brief RCOPE) described by Pargament et al., (2011). We used the Spanish version validated in the Puerto Rican population by Colón-Rivera (2014). The inventory measures PRC and NRC strategies based on the Pargament (1997) theoretical model. The instructions of the Brief-RCOPE invite participants to think about the most stressful event they have experienced in the last year. Then, the scale presents a list of 14 RC strategies (e.g., *I looked for God’s love and care; I looked for God’s help to release my courage; I wondered if God had abandoned me*), and ask the respondent to indicate, on a four-point Likert-type response scale, the degree to which each strategy applied to them: 1 (*Not at all*), 2 (*Somewhat*), 3 (*Quite a bit*) and 4 (*Very much*).

**Data analyses**

The IBM SPSS version 27.0 program (IBM Corp., 2020) was used to perform most statistical analyses. Descriptive statistics were calculated through measures of central tendency (mean, mode and median) and use of percent and frequencies, to explore the sociodemographic characteristics of the sample. In addition, the items’ discrimination index through corrected item-total correlation ($r_{bis}$) were considered. Those items with correlations magnitudes greater than .30 had acceptable discrimination indexes (Kline, 2005). The reliability of the measure was explored using the Cronbach’s alpha and the McDonald’s omega coefficients, both had to be equal or greater than .70 to be considered adequate (DeVellis, 2017). In addition, the convergent validity of the Brief RCOPE was examined through the average variance extracted (AVE) as recommended by Fornell and Larcker (1981). To establish convergent validity, the AVE had to be equal to or greater than .50, thus establishing that 50% or more of the construct’s variance was due to its indicators (Fornell & Bookstein, 1982). Concurrent validity was examined through a Pearson correlation coefficient between the PRC subscale and ratings in areas such as *importance of religion, participation in religious activities and participation in private religious practices* (e.g., prayer, sacred texts readings) which were extracted from the sociodemographic data form. For the examination of concurrent validity, Pearson correlation values less than .35 were considered weak or low correlations; values between .36 and .67 were considered moderate correlations; values between .68 and .89 were considered high correlations, and, finally, values from .90 onwards were considered very high correlations (Taylor, 1990). Finally, to determine the discriminant validity of each dimension, the value obtained by the individual AVE of each factor had to be higher than the maximum shared variance (MSV) and the average shared variance (ASV).

Using STATA version 15 program (StataCorp, 2017), two CFAs were conducted with the robust maximum likelihood estimation method. Specifically, the Satorra-Bentler adjustments were employed, which is a recommended alternative when data is not normally distributed (Satorra & Bentler, 2001), as it is the case in the current study’s measurement. In order to examine how the proposed model adjusted to the data, the following assessments were conducted: the corrected Chi-square test ($\chi^2_{sb}$), the ratio between
the latter and the degrees of freedom (χ²_{sb} / df), the corrected root mean square error of approximation (RMSEA_{sb}), the standardized root mean square residual (SRMR), the Tucker-Lewis index (TLI), the comparative fit index (CFI), and the Akaike information criterion (AIC). Values of χ²_{sb} / df lower than 3.0 were indicative of a very good fit for the model, while values of 5.0 or below were considered acceptable. Values of RMSEA less than .08, and SRMR values less than .08 were indicative of an acceptable adjustment of the model (Hooper, Coughlan, & Mullen, 2008; Kline, 2011). Meanwhile, CFI and TLI values greater than .90 represented acceptable adjustment of the model (Hooper et al., 2008; Kline, 2011). In addition, we used the AIC to compare the models’ parsimony. The model with the lower index shows a better adjustment (Schumacker & Lomax, 2010).

For the purpose of examining the statistical significance of the changes in the fit of the models when comparing one to the other, the Δ χ²_{sb} test was conducted (also known as the Satorra-Bentler scaled Chi-square difference test) with a p value of .05. So, as to evaluate the magnitude or size of such changes, Cohen’s w (Cohen, 1988) was employed, a coefficient whose suggested standards for a small (.1), medium (.3), and large (.5) effect are provided in parenthesis.

Results

Assessment of the normality assumptions

Using SPSS 27.0 (IBM Corp., 2020), the univariate normality assumption was tested using the Kolmogorov-Smirnov and the Shapiro-Wilks test. In both cases, the tests provided evidence that none of the 14 items had a normal distribution (p < .001). Tests for multivariate normality conducted with STATA also yielded results that revealed violations to the normality assumption [Mardia mSkewness = 62.61, χ²_{(560)} = 3186.91, p < .001; Mardia mKurtosis = 325.98, χ²_{(1)} = 1752.54, p < .001; Doornik-Hansen test, χ²_{(28)} = 2443.01, p < .001). Given the lack of normality of data, and to correct its effect on the estimation of the standard errors of parameters and global model fit, we used the Satorra-Bentler adjustments as part of the maximum likelihood estimation in STATA.

Confirmatory factor analyses

To determine the factor structure of the Brief RCOPE (Spanish version), two CFA were performed using the robust maximum likelihood estimation method. The first model evaluated was the one-dimensional model, in which the 14 original items were loaded onto one factor (M1). The CFA showed that the one-factor structure did not obtain adequate goodness of fit indexes. Then, a second model (M2) was examined with a two-correlated factor structure, in which items 1, 2, 3, 4, 5, 6 and 7 loaded on a common factor identified in the literature as PRC, and the items 8, 9, 10, 11, 12, 13 and 14 loaded on a common factor identified in the literature as NRC. Although this model was superior to the one-factor model (Table 2), item 14 yielded low factor loading (< .40) in the NRC latent variable and affected the goodness of fit statistics of the model. Therefore, the model was re-specified (see Figure 1) by eliminating this item (M2a). At this stage, the CFA revealed that the revised two-factor model of the Brief RCOPE provided the best adjustment to the data [Corrected χ² = 142.94, p < .001; Corrected RMSEA = .06; SRMR = .07; Corrected CFI = .96, Corrected TLI = .95; AIC = 7966.14 (see Table 2)]. Observed scores for the 7-item NRC factor and the 6-item version correlated at .99.
test yielded a \( p \) value of less than .001. Cohen’s \( w \) values for the M1 to M2 and the M2 to M2a comparison, which examined the size of the change in the \( \chi^2_{sb} \) value considering the change in degrees of freedom, were 1.30 (large size) and .12 (small size), respectively. This suggests that the differences in the fit indexes between the models examined were not trivial, although were substantially greater from M1 to M2.

### Table 2
Goodness-of-fit tests for analyzed models using robust maximum likelihood estimation.

| Model | \( \chi^2_{sb} \) | \( \chi^2_{sb}/df \) | RMSEA_{sb} | SRMR | CFI_{sb} | TLI_{sb} | AIC | \( \Delta \chi^2_{sb}(\Delta df) \) |
|-------|-----------------|----------------------|-------------|-------|---------|---------|-----|-----------------|
| M1    | 701.47          | 9.11                 | .16         | .20   | .70     | .64     | 9279.57 |                 |
| M2    | 194.93          | 2.57                 | .07         | .10   | .94     | .93     | 8517.41 | \( \Delta \chi^2_{sb}(1) = 70.80 \) |
| M2a   | 142.94          | 2.23                 | .06         | .07   | .96     | .95     | 7966.14 | \( \Delta \chi^2_{sb}(12) = 50.25 \) |

**Note.** Degrees of freedom for M1, M2 and M2a are 77, 76 and 64, respectively. M1 = one-dimensional model with 14 items; M2 = two-correlated-factors model; M2a = two-correlated-factors model with 13 items (deleting item 14); sb = Satorra-Bentler adjustments; \( \chi^2_{sb} \) = corrected Chi-square Test; df = degrees of freedom; RMSEA_{sb} = corrected root mean square error of approximation; SRMR = standardized root mean square residual; CFI_{sb} = corrected comparative fit index; TLI_{sb} = corrected Tucker-Lewis index; AIC = Akaike information criterion; \( \Delta \chi^2_{sb} \) = Satorra-Bentler scaled Chi-square difference test; \( \chi^2_{sb} \) values and \( \Delta \chi^2_{sb} \) tests are significant at \( p < .001 \).
All the Brief RCOPE items obtained discrimination indexes greater than .30 using the corrected item-total correlation technique as recommended (Kline, 2005). Table 4 shows the discrimination indexes of all items. In terms of the Cronbach’s reliability, the PRC subscale obtained an excellent coefficient of .94 and the NRC subscale showed an alpha value of .84. The omega coefficient was .94 for the PRC subscale and .85 for the NRC subscale. The convergent and discriminant validity of the revised two-correlated-factor model was also examined through the AVE, ASV, and MSV. Results showed that the AVE values for both factors were higher than the values of MSV and ASV (see Table 3). Furthermore, the PRC subscale correlated positively and significantly with the importance toward religious belief ($r = .681, p < .001$), participation in religious activities ($r = .569, p < .001$), and participation of private religious practices ($r = .568, p < .001$). However, the NRC did not correlate with these variables. The results showed that the Brief-RCOPE has a good convergent and discriminatory validity.

**Reliability and Validity of the Brief-RCOPE**

The present study aimed to examine the psychometric properties and factor structure of the Brief-RCOPE. The CFA showed a satisfactory fit with the data to the bifactorial structure of the Brief-RCOPE, particularly the model. These results are consistent with other studies conducted in Spanish-speaking countries, such as Chile (García et al., 2021), Argentina (Mezzadra & Simkin, 2017) and Mexico (Martinez & Sousa, 2011; Rivera-Ledesma & Montero-López, 2007), in which the Brief-RCOPE obtained a bifactor structure and good reliability scores. For the revised two-factor model, item 14 was removed given the improvement in the goodness-of-fit of the structure model associated with its exclusion. The relative weakness of item 14 for our sample is similar to findings from Freitas et al. (2015). It should be noted that retaining the item in the NRC would also reduce the alpha coefficient of the factor to .84, the omega coefficient to .84, and the AVE to .451. This latter value would be below the requested level (of .50 or more) to support the convergent validity of the NRC factor, and is considered unacceptable. However, the overall results of the revised model replicate the two-dimensional structure considered by the authors in the theoretical construction of the instrument. The two-factor structure of the Brief-RCOPE is closely related to the theoretical foundations and assumptions of the coping model. Furthermore, this instrument obtained adequate Cronbach alpha internal consistency and omega coefficient for the PRC and NRC dimensions, which is cont-

### Table 3

| Factors                  | AVE  | MSV  | ASV  | Factor 1 | Factor 2 |
|--------------------------|------|------|------|----------|----------|
| Positive Religious Coping| .69  | .03  | .03  | 1        | .23***   |
| Negative Religious Coping| .51  | .03  | .03  | .17**    | 1        |

**Note.** The value below the diagonal represents the correlation between latent factors, while the value above the diagonal represent the correlation among direct scores AVE = average variance extracted; MSV = maximum shared variance; ASV = average shared variance; Brief R-COPE = Brief Religious Coping Scale. **$p < .01$; ***$p < .001$. 

**Discussion**

The present study aimed to examine the psychometric properties and factor structure of
### Table 4
Item discrimination indexes and confidence intervals for factor loadings.

| Items of the Brief R-COPE (in Spanish) | $r_{bis1}$ | $r_{bis2}$ | $\beta$ | 95% CI$_{sb}$ |
|----------------------------------------|------------|------------|---------|--------------|
| 1. Busqué una conexión más fuerte con Dios. | .87 | .92 | .90 – .95 |
| 2. Busqué el amor y cuidado de Dios. | .90 | .94 | .93 – .96 |
| 3. Busqué ayuda de Dios para soltar mi coraje. | .80 | .83 | .79 – .88 |
| 4. Intenté resolver la situación de la mano de Dios. | .87 | .90 | .88 – .93 |
| 5. Trató de ver cómo Dios podría estar tratando de fortalecerme en esta situación. | .80 | .82 | .78 – .87 |
| 6. Pedi perdón por mis pecados. | .71 | .72 | .66 – .78 |
| 7. Me enfoqué en la religión para dejar de preocuparme por mis problemas. | .61 | .63 | .57 – .69 |
| 8. Me pregunté si Dios me había abandonado. | | | | |
| 9. Sentí que Dios me había castigado por mi falta de devoción (consagración o fervor). | .71 | .81 | .74 – .89 |
| 10. Me pregunté qué hice para que Dios me castigara así. | .80 | .90 | .86 – .95 |
| 11. Dudé del amor de Dios por mí. | .66 | .72 | .62 – .81 |
| 12. Me pregunté si mi iglesia me había abandonado. | .43 | .43 | .29 – .57 |
| 13. Decidí que el diablo (Satanás, Lucifer o el mal) había hecho que esto sucediera. | .45 | .46 | .34 – .59 |

Note. N = 302. $r_{bis1}$ = corrected item-total correlations of items with the positive religious coping factor; $r_{bis2}$ = corrected item-total correlations of items with the negative religious coping factor (revised two-factor model); $\beta$ = standardized regression coefficient for each item with its respective factor in the revised two-factor model; CI$_{sb}$ = confidence interval with the Satorra-Bentler correction for non-normality; Brief R-COPE = Brief Religious Coping Scale. All coefficients are statistically significant at $p < .001$.

Consistent with previous psychometric data about the scale when used with Puerto Ricans (Colón-Rivera, 2014). All the items obtained an adequate discrimination index. The AVE, MSV and ASV of the subscales were excellent, showing a good concurrent, convergent, and discriminant validity. The data suggest that the Brief-RCOPE is a reliable and valid instrument to measure RC strategies among Puerto Rican adults, particularly if item 14 is excluded.

Despite the fact there is a religious coping measure validated with Puerto Rican samples (González-Rivera & Pagán-Torres, 2018), the measure is based in the Lazarus and Folkman (1986) theoretical model of external and internal coping strategies. However, the Brief RCOPE is based on the Pargament theoretical model of positive and negative religious coping strategies (Pargament, 1997). In fact, this is the original model that conceptualized the religious coping strategies as a variable of study. Furthermore, the Brief-RCOPE is a measure widely used in a diversity of countries to measure RC. Therefore, the examination of its reliability, validity and factor structure is essential to promoting the scientific study of religion and spirituality in Puerto Rico, as well as to comparing the findings from studies conducted in other countries to the results ob-
The Brief-RCOPE may be used in a clinical and research context with clinical and non-clinical samples. In the clinical setting, this measure may be used as a tool for screening religious strategies with high precision and in a short time. Another advantage of this measure is that it provides the opportunity to simultaneously administer a battery of other religious/spiritual and mental health measures (for clinical or research purposes) to explore additional constructs that are positively and negatively associated with RC. The findings provide preliminary evidence of the validity and factor structure of the Brief-RCOPE in the Puerto Rican context. This study adds support to the relevance of conducting additional research in Puerto Rico aimed to evaluate the relationship that PRC and NRC dimensions may have with mental health variables in clinical and non-clinical samples.

Our study has several limitations. First, the participants in the study were not randomly recruited. Instead, we used a non-probabilistic convenience recruitment method. Therefore, our sample is not representative of the Puerto Rican adult population. Second, we did not evaluate the reliability of the instrument over time (using a test-retest strategy). However, we did evaluate Cronbach alpha consistency and McDonald’s omega coefficient. In addition, digital recruitment allowed us to amplify the diversity of the sample in terms of sociodemographic characteristics and sample size. Moreover, we used advanced statistical techniques with CFA and an adequate sample size to provide empirical strength to our results. Third, the number of women participants was significantly higher than the number of men participants in this study. Further studies should consider exploring factorial invariance and external validity (correlations with other psychological variables), among other relevant analyses. In addition, future research should consider working with religious samples, as working with university students or the general population is one of the main limitations within the field of Psychology of Religion and Spirituality (Kapuscinski & Masters, 2010). Despite the limitations mentioned above, the results of this study provide relevant and preliminary information on the psychometric properties and factor structure of the Brief RCOPE in Puerto Rican adults. In addition, we suggest future research to focus on using a diversity of psychological measures that have been positively and negatively associated with RC to examine the nature of PRC and NRC strategies as potential protective or risk factors in diverse populations.

**Conclusion**

This is the first study aimed to examine the psychometric properties and factor structure of the Brief-RCOPE in Puerto Ricans. Our findings revealed that the Spanish Brief-RCOPE used in Puerto Rico has excellent psychometric properties and essentially replicates the two-dimensional factor structure. These findings support the applicability of the Brief-RCOPE within the Puerto Rican population. Future studies could further explore the relevance of Brief-RCOPE dimensions with other religious/spiritual measures and mental health outcomes in Puerto Rican clinical (outpatient) and non-clinical samples. This research represents a significant contribution to the scientific study of religion and spirituality in Puerto Rico. In summary, the Brief-RCOPE is a reliable and valid measure, easy to administer, that may be used in any research and/or clinical setting to explore religious coping strategies and their potential status as protective or risk factors on mental health.
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