On an imaginary exponential functional of Brownian motion

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Abstract

We investigate a random integral which provides a natural example of an imaginary exponential functional of Brownian motion. This functional shows up in the study of the binary annihilation process, within the Doi–Peliti formalism for reaction-diffusion systems. The main emphasis is put on the complementarity between the usual Langevin approach and another approach based on the similarity with Kesten variables and other one-dimensional disordered systems. Even though neither of these routes leads to the full solution of the problem, we have obtained a collection of results describing various regimes of interest.

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(Some figures in this article are in colour only in the electronic version)

1. Introduction

Real exponential functionals of Brownian motion have been the subject of much activity in probability theory [1]. They have found many applications to problems ranging from finance to physics [2]. To be more specific, if \( B(t) \) is a Brownian motion (a Wiener process with \( B(0) = 0 \) and \( \langle B(t)^2 \rangle = t \)), the following stochastic variable:

\[
X(t) = \int_0^t e^{-s+gB(s)} \, ds,
\]

(1.1)

where \( g \) is a real coupling constant, shows up in one guise or another in various models of disordered systems [3, 4]. The above random integral represents the solution with the initial value \( X(0) = 0 \) of the following Langevin equation with multiplicative noise:

\[
\frac{dX(t)}{dt} = 1 - X(t) + g X(t) \eta(t),
\]

(1.2)
with \( \eta(t) = dB(t)/dt \) being a zero-mean Gaussian white noise, normalized as
\[
\langle \eta(s)\eta(t) \rangle = \delta(t-s).
\] (1.3)

The derivation of the solution to (1.2) in the form (1.1), including the required stochastic calculus prescription (Stratonovich), will be reviewed in section 2. As an illustration of the versatility of the situations in which the random variable \( X(t) \) appears, the Langevin equation (1.2) may serve to study the effect of multiplicative noise on a deterministic fixed point. Indeed, the linear force \( 1 - X(t) \) stabilizes \( X(t) \) at the fixed point \( X = 1 \) in the absence of noise. The solution \( X(t) \) however keeps on fluctuating forever because of the multiplicative noise term \( gX(t)\eta(t) \). In the long-time limit, what emerges out of this interplay between the deterministic restoring force and the fluctuating noise term is a non-trivial distribution for the random variable
\[
X = \int_0^\infty e^{-t+gB(t)} \, dt,
\] (1.4)

representing the stationary solution of (1.2). The distribution of \( X \equiv \lim_{t \to \infty} X(t) \) can be computed in a number of ways, and we shall review in section 2 two useful methods to do so. Using either more elaborate path-integral methods or probabilistic identities, the full time-dependent distribution of \( X(t) \) can also be obtained and features, among other things, an interesting continuous spectrum of relaxation rates [2].

A seemingly innocuous and somewhat natural generalization of (1.4) consists in analytically continuing the coupling constant \( g \to ig \). We thus obtain the following random integral, defining an imaginary exponential functional of Brownian motion:
\[
Z = \int_0^\infty e^{-t+igB(t)} \, dt,
\] (1.5)
which now lives in the complex plane (actually in the unit disk).

The main goal of this paper is to investigate the distribution of the functional \( Z \). Throughout the following, instead of the coupling constant \( g \), we rather make use of the dimensionless diffusion constant
\[
D = \frac{g^2}{2}.
\] (1.6)

As hinted above, even though a large number of works have been devoted to real exponential functionals of Brownian motion, much less is known about the distribution of complex functionals of Brownian motion such as (1.5). We found it interesting from a conceptual viewpoint to tackle this problem, particularly in order to see if any of the methods which proved successful for the real case would extend to this complex-variable setting. Besides this, it turns out that complex stochastic processes have surfaced time and again in different scientific disciplines ranging from signal theory, where processes involving imaginary exponential functionals of Brownian motion occur in the study of phase noise [5], to quantum optics [6], in conjunction with the development of phase-space representations and the associated formalism of quasi-probabilities, and finally to reaction–diffusion processes [7, 8] through the Doi–Peliti approach [9]. The latter topic has constituted our original thrust to embark on the study of (1.5). Let us now describe how the connection emerges.

In the context of interacting particle systems such as reaction–diffusion systems, there exists a standard set of techniques, usually referred to as the Doi–Peliti formalism [9] (for a recent review, see [10] and references therein), which allows us to recast the master equation describing the evolution of these processes in terms of a field theory whose action involves

\[^3\] Here and throughout the following the symbol \( \equiv \) denotes an equality in law (i.e. in distribution).
a pair of conjugate fields. Without entering into much detail, provided certain technical conditions are met, the theory can in turn be transformed into a Langevin equation for a single density field \( \phi_n(t) \), customarily dubbed so because its noise-average \( \langle \phi_n(t) \rangle \) coincides with the local mean particle number for the underlying reaction-diffusion process \( \langle \rho_n(t) \rangle \). Here the double brackets denote an average over the dynamics of the particles. In the case of the binary annihilation reaction
\[ A + A \rightarrow \emptyset, \] (1.7)
where particles \( A \) diffuse by hopping with rate \( \kappa \) on a hypercubic lattice in dimension \( d \) and annihilate pairwise with rate \( \lambda \) when they meet on a given lattice site \( n \), one can show [7–9] that the stochastic density field \( \phi_n(t) \) obeys the following Langevin–Itô equation:
\[
\frac{d\phi_n(t)}{dt} = \kappa (\nabla^2 \phi)_n(t) - \lambda \phi_n(t)^2 + \zeta_n(t),
\] (1.8)
with \( \nabla^2 \) being the lattice Laplacian, such that, e.g., \( (\nabla^2 \phi)_n = \phi_{n+1} + \phi_{n-1} - 2\phi_n \) in one dimension. It can be expected on physical grounds that the amplitude of the Gaussian noise term \( \zeta_n[\phi_n] \) vanishes when \( \phi_n = 0 \). Indeed, the first two terms on the right-hand side of (1.8) respectively account for the diffusion of the particles and (mean-field) decay rate of the particle density due to pairwise annihilation. If the noise amplitude vanishes when \( \phi_n = 0 \), there is no evolution at all in regions where \( \phi_n = 0 \). The Doi–Peliti approach also yields the following expression for the correlator of the noise:
\[
\langle \zeta_m(s)\zeta_n(t) \rangle = -\lambda \phi_n(t)^2 \delta_{m,n} \delta(t - s),
\] (1.9)
which therefore has a negative variance. In other words, we have \( \zeta_n(t) = i\sqrt{\lambda} \phi_n(t) \eta_n(t) \), where \( \eta_n(t) \) is a normalized real Gaussian white noise. This explains why (1.8) is often referred to as an imaginary-noise equation. One would obtain the same Langevin equation using Gardiner’s quasi-probability formalism, where the particle-number probability distribution is represented as a superposition of Poisson distributions with weights \( \phi_n \) [8, 11]. The equivalence between the Doi–Peliti formalism and Gardiner’s Poisson representation method has been demonstrated in general in [12]. There is no contradiction in either formalism, as soon as the auxiliary field \( \phi_n(t) \) is complex-valued, provided one refrains from erroneously identifying \( \phi_n(t) \) with the (integer-valued) stochastic variable \( \rho_n(t) \), based on the sole equality between the mean values \( \langle \phi_n(t) \rangle = \langle \rho_n(t) \rangle \). In fact, the precise relationship between the distribution of \( \phi_n(t) \) and that of \( \rho_n(t) \) is that the ordinary moments of \( \phi_n(t) \) are equal to the factorial moments of \( \rho_n(t) \). This is a particular instance of what is referred to as duality between two stochastic processes in the probabilistic literature [13]. Thus in particular
\[
\langle \rho_n(t)\rho_n(t - 1) \rangle = \langle \phi_n(t)^2 \rangle,
\] (1.10)
and so one can have \( \langle \phi_n(t)^2 \rangle < 0 \), while keeping \( \operatorname{var} \rho_n(t) = \langle \rho_n(t)^2 \rangle - \langle \rho_n(t) \rangle_2 = \langle \phi_n(t)^2 \rangle + \langle \phi_n(t) \rangle - \langle \phi_n(t) \rangle_2 \geq 0 \). It is also commonly accepted that the complex-valued nature of the trajectories of the field \( \phi_n(t) \) is needed in order to account for the importance of fluctuation effects in low spatial dimensions, resulting in a slower decay for the total density of particles than what the naive law of mass action would predict (namely \( t^{-d/2} \) versus \( 1/t \) in \( d < 2 \) for reaction (1.7) [10, 14]).

Henceforth, along the lines of [7, 8], we focus onto the single-site problem associated with (1.8), neglecting any spatial dependence. This simplification will allow us to better understand the role of the excursions of the field \( \phi(t) \) in the complex plane. Setting \( \kappa = 0 \), and absorbing the reaction rate \( \lambda \) into the time scale, one ends up with the following Langevin–Itô equation for a single complex stochastic variable \( \psi(t) \):
\[
\frac{d\psi(t)}{dt} = -\psi(t)^2 + \psi(t) \eta(t),
\] (1.11)
where $\eta(t)$ is a normalized Gaussian white noise (see (1.3)). The initial condition $\varphi(0)$ is real and non-negative (e.g. $\varphi(0) = \rho_0$ if one starts from a Poisson distribution with density $\rho_0$ for the original particle system). The conjugation symmetry $\varphi \to \bar{\varphi}$ ensures that the imaginary part of $\varphi(t)$ averages over to zero, so as to maintain the reality and the non-negativity of $\langle \varphi(t) \rangle = \langle \rho(t) \rangle$. It has already been noticed by several authors [6–8] that (1.11) becomes linear in the variable $Z(t) = 1/\varphi(t)$. One thus obtains the explicit solution

$$Z(t) = Z(0) e^{-t/2} + \int_0^t e^{-\frac{(t-s)}{2} + \frac{i}{2}|B(s) - B(t)|} \, ds. \quad (1.12)$$

Rescaling time and using the scaling property $B(at) = \sqrt{a} B(t)$ of Brownian motion, we obtain by identifying (1.5) and (1.12) the following identity between the stationary solution $Z \equiv \lim_{t \to \infty} Z(t)$ and the functional $Z$ for $g = \sqrt{2}$, i.e. $D = 1$ [8]:

$$Z = \lim_{t \to \infty} Z(t) \equiv 2Z|_{D=1} = 2 \int_0^\infty e^{-t + i\sqrt{2}B(t)} \, dt. \quad (1.13)$$

The above representation of the stationary solution has striking consequences. At the level of the original process (1.7), the stationary state is rather featureless, as there just remains either zero or one particle, depending on the parity of the initial condition. By (1.10), and the corresponding equations for higher-order moments, this implies that $\langle \varphi^p \rangle = 0$ for any integer $p \geq 2$. Owing to (1.13), this reads $\langle Z^{-p} \rangle|_{D=1} = 0$. The full distribution of $Z$ will however turn out to be highly non-trivial. In particular, in stark contrast to what intuition backing up (1.8) or (1.11) could let us foresee, in the stationary state the variable $\varphi = 2/Z$ never reaches the value $\varphi = 0$, characteristic of the absorbing state. Indeed, as already announced, $Z$ lies within the unit disk.

The setup of this paper is as follows. In section 2, we present a self-contained investigation of the real exponential functional $X$ (see (1.4)). The emphasis is put on the complementarity between the usual Langevin approach and another approach based on the similarity with the random recursions met in the study of Kesten variables and one-dimensional disordered systems. The main section of the paper (section 3) is devoted to a detailed study of the imaginary exponential functional $Z$ (see (1.5)). We present numerical illustrations of the distribution of $Z$ and investigate many facets of the problem by analytical means, including the relationship between the Langevin and Kesten approaches, the moments $\langle Z^k \bar{Z}^l \rangle$, the weak-disorder and strong-disorder regimes, and the asymptotic behavior of the distribution near the unit circle. Section 4 contains a brief summary of our findings.

2. A warming up: the real functional

This section is to a large extent intended as a warming up. It is devoted to a self-contained study of the real exponential functional (see (1.4))

$$X = \int_0^\infty e^{-t + g\sqrt{2}B(t)} \, dt. \quad (2.1)$$

The positive random variable thus defined is one of the exponential functionals of the Brownian motion which have been investigated in probability theory, chiefly by Yor and his collaborators (see [1] for a review). It also appears in the physics literature, in the context of one-dimensional disordered systems [2–4]. The coupling constant $g$ measuring the strength of noise or, equivalently, the diffusion constant $D$ (see (1.6)) is the sole parameter entering the definition of $X$. 
2.1. Langevin approach

A first approach to study the random variable $X$ consists in using Langevin equations. At this point it is useful to recall some elements of stochastic calculus. A Langevin equation of the form

$$\frac{dX(t)}{dt} = a(X(t)) + b(X(t))\eta(t)$$  \hspace{1cm} (2.2)

is ambiguous as soon as it is nonlinear, in the sense that the noise $\eta(t)$ multiplies a non-trivial function $b(X(t))$ of the position $X(t)$. This ambiguity is due to the usage of a continuous-time formalism, and can be lifted in many ways. Although this is a well-known topic, which is discussed at length in various books (see e.g. [11, 15–17]), we nevertheless found it useful to review the so-called Stratonovich versus Itô dilemma, as we shall constantly switch back and forth between these two viewpoints.

- **Stratonovich prescription.** The Langevin–Stratonovich differential equation

$$\left[S\right] \frac{dX(t)}{dt} = a_S(X(t)) + b(X(t))\eta(t)$$ \hspace{1cm} (2.3)

can be essentially thought of as an ordinary differential equation. It is amenable to nonlinear changes of variable according to the usual rules of differential and integral calculus. The main disadvantage is that $X(t)$ and $\eta(t)$ at the same time $t$ are not independent.

- **Itô prescription.** The Langevin–Itô differential equation

$$\left[I\right] \frac{dX(t)}{dt} = a_I(X(t)) + b(X(t))\eta(t)$$ \hspace{1cm} (2.4)

has the advantage that the process $X(t)$ and the noise $\eta(t)$ at the same time $t$ are independent, so that one has e.g. $\langle a_I(X) \rangle = 0$ in the stationary state. Equation (2.4) also provides a natural discretization of the process $X(t)$. Considering discrete times $t = n \varepsilon$ so that $X_n \equiv X(t_n)$, we obtain the recursion

$$X_{n+1} = X_n + a_I(X_n)\varepsilon + b(X_n)\zeta_{n+1},$$ \hspace{1cm} (2.5)

where $\zeta_{n+1}$ is a Gaussian random variable, independent of $X_n$, such that $\langle \zeta_{n+1} \rangle = 0$ and $\langle \zeta_{n+1}^2 \rangle = \varepsilon$. This discrete scheme can be efficiently used in a numerical simulation. The main disadvantage of the Itô prescription is that care must be exercised when making nonlinear changes of variable.

Both Langevin equations (2.3) and (2.4) describe the same stochastic process $\{X(t)\}$ if their drift terms are related to each other by the correspondence formula

$$a_I(X) - a_S(X) = \frac{1}{2} \frac{d}{dX} b(X)^2 = \frac{1}{2} b(X) \frac{db(X)}{dX}. \hspace{1cm} (2.6)$$

The corresponding time-dependent probability density $P(x, t) = \langle \delta(X(t) - x) \rangle$ obeys the (forward) Fokker–Planck equation

$$\frac{\partial P}{\partial t} = - \frac{\partial}{\partial x} (a_I(x)P) + \frac{1}{2} \frac{\partial^2}{\partial x^2} (b(x)^2 P)$$

$$= - \frac{\partial}{\partial x} (a_S(x)P) + \frac{1}{2} \frac{\partial}{\partial x} \left( b(x) \frac{\partial}{\partial x} (b(x) P) \right). \hspace{1cm} (2.7)$$

The independence of the observable pdf $P(X, t)$ from the underlying stochastic calculus prescription is precisely guaranteed by the correspondence formula (2.6). Demanding from
either of the above formulas that the probability current be zero determines the stationary density of the process, 
\[ f_X(x) = \lim_{t \to \infty} P(x, t), \]
according to the two equivalent expressions:
\[ f_X(x) = N_I \exp \left( 2 \int_{x_0}^x \frac{a_I(y)}{b(y)^2} \, dy \right) = N_S \exp \left( 2 \int_{x_0}^x \frac{a_S(y)}{b(y)^2} \, dy \right), \]  
(2.8)
where \( x_0 \) is an arbitrary initial point, and \( N_I \) and \( N_S \) are normalization constants.

It is now time to return to our functional \( X \) (see (1.4)). The Langevin–Stratonovich equation (1.2), i.e.
\[ [S] \frac{dX(t)}{dt} = 1 - X(t) + gX(t)\eta(t), \]  
(2.9)
is equivalent to the Langevin–Itô equation
\[ [I] \frac{dX(t)}{dt} = 1 + (D - 1)X(t) + gX(t)\eta(t). \]  
(2.10)
Indeed \( b(X) = gX \) and \( a_S(X) = 1 - X \) yield \( a_I(X) = 1 + (D - 1)X \) (see (1.6)).

The Langevin–Stratonovich equation (2.9) can be readily integrated. We thus obtain the explicit stochastic representation of the process as
\[ X(t) = X(0) e^{-t+gB(t)} + \int_0^t e^{-(t-s)+g(B(t)-B(s))} \, ds. \]  
(2.11)
In the \( t \to \infty \) limit, the above expression loses the memory of its initial condition \( X(0) \). Moreover, using the stationarity of the Brownian motion, the integral can be recast as an integral over \( \tau = t - s \), which identifies with (2.1) or (1.4). We have thus shown that the exponential functional \( X \) represents the stationary solution of (2.9) or (2.10). In other words we have \( X \equiv \lim_{t \to \infty} X(t) \). In particular \( X(t) \) is positive in the \( t \to \infty \) limit, irrespective of the initial value \( X(0) \).

The Itô prescription allows us to directly read off the stationary mean value of \( X \) from (2.10):
\[ \langle X \rangle = \frac{1}{1 - D}. \]  
(2.12)
This expression only makes sense for \( D < 1 \). It is in agreement with the behavior of the time-dependent mean value \( \langle X(t) \rangle \). Equation (2.10) indeed yields
\[ \frac{d\langle X(t) \rangle}{dt} = 1 + (D - 1)\langle X(t) \rangle, \]  
(2.13)
whose solution, i.e.
\[ \langle X(t) \rangle = \frac{1}{1 - D} + \left( X(0) - \frac{1}{1 - D} \right) e^{(D-1)t}, \]  
(2.14)
relaxes exponentially fast to (2.12) for \( D < 1 \), whereas it diverges for \( D > 1 \). Result (2.14) can be recovered by averaging (2.11) over the Brownian motion \( \{B(s)\} \).

Equation (2.8) leads to the following explicit expression for the distribution of the functional \( X \) [1, 3, 4]:
\[ f_X(x) = \frac{1}{D^{1/D} \Gamma(1/D)} x^{-(1+1/D)} e^{-1/(DX)}. \]  
(2.15)
In other words, the variable \( Y = 1/(DX) \) is distributed according to a gamma law with the parameter \( 1/D \):
\[ f_Y(y) = \frac{1}{\Gamma(1/D)} y^{-(1+1/D)} e^{-y}. \]  
(2.16)
Expression (2.15) falls off exponentially fast as \( x \to 0 \). It exhibits a fat tail at large \( x \), as it decreases as a power law, with the continuously variable exponent \(-\frac{1 + 1}{D}\). As a consequence, the moment \( \langle X^s \rangle \) only converges for \( s < \frac{1}{D} \) (see (2.31)). This provides another way to explain the divergence of (2.12) at \( D = 1 \).

Finally, at par with the qualitative discussion given in the Introduction, one can interpret the emergence of a fat tail at large values of \( X \) and the possible divergence of the mean value \( \langle X \rangle \) as consequences of the exchange of stability between the two fixed points of the dynamical system (2.9), i.e. that at \( X = 1 \) which is stable for the deterministic dynamics, i.e. in the absence of noise, and a fluctuating one sitting formally at \( X = \infty \).

Figure 1 shows a plot of the density \( f_X \) for three values of \( D \).

2.2. Kesten approach

An alternative approach to study the random variable \( X \) consists in using the similarity between the random integral (2.1) and random sums of products referred to as Kesten variables. Let us start with a reminder. A Kesten variable \([18–22]\) is defined as

\[
Z = 1 + \xi_1 + \xi_1 \xi_2 + \xi_1 \xi_2 \xi_3 + \cdots ,
\]

(2.17)

where the \( \xi_n \) are i.i.d. positive random variables with probability density \( f_\xi (\xi) \). If the latter distribution is such that \( \langle \ln \xi \rangle < 0 \), sum (2.17) is almost surely convergent, and it represents the stationary solution of the random recursion

\[
Z_{n+1} = 1 + \xi_{n+1} Z_n,
\]

(2.18)

where \( \xi_{n+1} \) is independent of \( Z_n \) and has density \( f_\xi (\xi) \). In other words, we have the identity among random variables

\[
Z \equiv 1 + \xi Z',
\]

(2.19)

where \( Z' \) is a copy of \( Z \), and \( \xi \) is independent of \( Z' \). The density \( f_Z (z) \) of the Kesten variable \( Z \) obeys the integral equation

\[
f_Z (z) = \int_0^\infty \frac{d\xi}{\xi} f_\xi (\xi) f_Z \left( \frac{z-1}{\xi} \right),
\]

(2.20)

which cannot be solved in closed form in general.
Let $\xi_{\text{min}} = a$ and $\xi_{\text{max}} = b$ be the smallest and largest values of $\xi$ (lower and upper bounds of the support of $f(\xi)$), and similarly $Z_{\text{min}} = A$ and $Z_{\text{max}} = B$ the smallest and largest values of $Z$. The condition $\ln \xi < 0$ implies $a < 1$. We have then $A = 1/(1 - a)$. If $a < b < 1$, we have $B = 1/(1 - b)$ and $f_2$ has finite support. In the more interesting situation where $a < 1 < b$, the support of $f_2$ extends to infinity. It is known that the distribution $f_2$ generically exhibits a fat tail, i.e. a power-law fall-off, of the form

$$f_Z(z) \sim z^{-(1+\alpha)}, \quad (2.21)$$

where the exponent $\alpha > 0$ is given by the condition $[18, 20, 21]

$$\langle \xi^\alpha \rangle = 1. \quad (2.22)$$

The density $f_Z(z)$ has been derived explicitly $[19, 21]$ in the case where $f_X(\xi)$ is a power law on an interval of the form $[0, b]$ or $[a, \infty]$.

The exponential functional $X$, defined in (2.1) as an integral of an exponential, can be viewed as a continuous analogue of the Kesten variable $Z$, defined in (2.17) as a sum of products. Similarly, the Langevin equation (2.9) is the continuous time analogue of recursion (2.18). This correspondence, already referred to in $[2, 3]$, can be understood quantitatively in the following way. Let us introduce a small time step $\epsilon$ and split the integral in (2.1) as $X = X^{(1)} + X^{(2)}$, where

$$X^{(1)} = \int_0^\epsilon e^{-\epsilon g B(t)} dt, \quad X^{(2)} = \int_\epsilon^\infty e^{-\epsilon g B(t)} dt. \quad (2.23)$$

To first order in $\epsilon$, we have $X^{(1)} = \epsilon$. Furthermore, setting $t = \epsilon + s$, we have $B(t) \equiv B(\epsilon) + B(s)$, so that $X^{(2)} = e^{-\epsilon g} X'$, where $X'$ is a copy of the variable $X$ and $\xi = g B(\epsilon)$ is a Gaussian variable, independent of $X'$, such that $\langle \xi^\alpha \rangle = 0$ and $\langle \xi^{1+\alpha} \rangle = g^2 = 2D\epsilon$. Putting everything together, the exponential functional $X$ appears (in the $\epsilon \to 0$ limit) to obey the identity

$$X \equiv \epsilon + \xi X', \quad (2.24)$$

with

$$\xi = e^{-\epsilon \xi}, \quad (2.25)$$

where $\xi$ is Gaussian, such that $\langle \xi \rangle = 0$ and $\langle \xi^2 \rangle = 2D\epsilon$. Up to an unimportant global factor $\epsilon$, the exponential functional $X$ therefore identifies (in the $\epsilon \to 0$ limit) with the Kesten variable $Z$ generated by the random input variables $\xi$ given by (2.25). This is precisely the Kesten variable investigated in $[22]$, where the main emphasis is already put on the $\epsilon \to 0$ limit, and where distribution (2.15) is derived in this limit.

The above correspondence directly yields the fall-off exponent of the density of $X$. We have indeed $\langle e^{\epsilon \xi} \rangle = e^{\epsilon^2 \langle \xi^2 \rangle/2} = e^{\epsilon^2 \xi}$, so that

$$\langle \xi^\epsilon \rangle = e^{\epsilon^2(Ds - 1)\epsilon}. \quad (2.26)$$

Condition (2.22) thus predicts $\alpha = 1/D$, in agreement with (2.15).

The full distribution of $X$ can actually be derived from identity (2.24). Consider indeed the moment function

$$M(s) = \langle X^s \rangle. \quad (2.27)$$

Equations (2.24) and (2.26) can be respectively expanded in the $\epsilon \to 0$ limit to yield

$$M(s) = \langle \xi^\epsilon \rangle M(s) + s \langle \xi^{\epsilon-1} \rangle M(s - 1)\epsilon + \cdots \quad (2.28)$$

and

$$\langle \xi^\epsilon \rangle = 1 + s(Ds - 1)\epsilon + \cdots, \quad (2.29)$$

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where the dots stand for terms of higher order in $\varepsilon$. We are left (in the $\varepsilon \to 0$ limit) with the functional equation

$$M(s - 1) = (1 - Ds)M(s),$$

(2.30)

whose normalized solution reads

$$M(s) = \frac{\Gamma(-s + 1/D)}{\Gamma(1/D)D^s} \quad (\text{Re} \, s < 1/D).$$

(2.31)

Setting $s = 1$, we recover expression (2.12) for $\langle X \rangle$, provided $D < 1$. The density of $X$ is given by the inverse Mellin transform

$$f_X(x) = \int \frac{ds}{2\pi i} x^{-s-1} M(s).$$

(2.32)

Expression (2.15) is recovered by summing the contributions of the poles of the integrand at $s = 1/D + n$ for $n = 0, 1, \ldots$. The leftmost pole at $s = 1/D$ is responsible for the power-law tail with the exponent $-(1 + 1/D)$.

3. The imaginary functional

We now turn to the main object of this paper, namely the distribution of the random variable $Z$ defined by the integral (1.5), i.e.

$$Z = \int_0^\infty e^{-t+igB(t)} \, dt.$$  

(3.1)

From a formal viewpoint, the complex variable $Z$ can be viewed as the analytical continuation as $g \to ig$ of its real counterpart $X$, investigated in section 2. This continuation amounts to changing the sign of the diffusion constant ($D \to -D$). It is however worth emphasizing that the study of $Z$ is far more difficult than that of $X$, as most of the usual tools which are fit to investigate real random variables cease to work in the case of a complex random variable.

Setting $Z = X + iY$, we are equivalently interested in the joint distribution $f(x, y)$ of the two correlated real random variables

$$X = \int_0^\infty e^{-t} \cos(gB(t)) \, dt, \quad Y = \int_0^\infty e^{-t} \sin(gB(t)) \, dt.$$  

(3.2)

Let us start with a few general facts. The complex random variable $Z$ lives inside the unit disk, since

$$|Z| \leq \int_0^\infty e^{-t} \, dt = 1.$$  

(3.3)

Let us anticipate that its distribution has a smooth density $f(x, y)$, whose support is the whole unit disk. Before we turn to more specific features, it is illustrative to first evaluate the mean $\langle Z \rangle$. Using the identity

$$\langle e^{igB(t)} \rangle = e^{-g^2/2} \langle B(t)^2 \rangle = e^{-Dt},$$

(3.4)

we obtain at once

$$\langle Z \rangle = \int_0^\infty e^{-t} \langle e^{igB(t)} \rangle \, dt = \int_0^\infty e^{-(D+1)t} \, dt,$$

(3.5)

i.e.

$$\langle Z \rangle = \frac{1}{D + 1}.$$  

(3.6)
so that
\[ \langle X \rangle = \frac{1}{D + 1}, \quad \langle Y \rangle = 0. \]  
(3.7)
The above results obey the conjugation symmetry: \( Z = X + iY \) and \( \bar{Z} = X - iY \) have the same law. Result (3.6) can be recovered by performing the analytical continuation \( D \rightarrow -D \) on (2.12). At variance with the latter result, expression (3.6) depends smoothly on the diffusion constant \( D \), decreasing from \( \langle Z \rangle = 1 \) in the \( D \rightarrow 0 \) limit to \( \langle Z \rangle \rightarrow 0 \) in the \( D \rightarrow \infty \) limit. These limiting regimes will be respectively investigated in sections 3.5 and 3.6, whereas the main features of the dependence of the distribution of \( Z \) on \( D \) will be studied in section 3.7.

3.1. Langevin approach

It can be shown along the lines of section 2.1 that the random variable \( Z \equiv \lim_{t \rightarrow \infty} Z(t) \) represents the stationary solution of the Langevin–Stratonovich equation
\[ [S] \quad \frac{dZ}{dt} = 1 - Z + ig\eta(t), \]  
(3.8)
where \( \eta(t) \) is again a Gaussian white noise. The solution to (3.8) reads
\[ Z(t) = Z(0)e^{-t+igB(t)} + \int_0^t e^{-(t-s)+ig(B(t)-B(s))} \, ds. \]  
(3.9)
A comparison with (1.12) shows that identity (1.13) extends to finite times as \( Z(t) \equiv 2Z(t/2) \) (again with \( D = 1 \)).

The Langevin–Itô equation corresponding to (3.8) is
\[ [I] \quad \frac{dZ}{dt} = 1 - (D + 1)Z + ig\eta(t). \]  
(3.10)
This equation yields \( \langle 1 - (D + 1)Z \rangle = 0 \) in the stationary state. We thus readily recover expression (3.6) of \( \langle Z \rangle \). Equation (3.10) also provides an efficient discrete scheme. Considering discrete times \( t = n\varepsilon \), and adapting recursion (2.5) to this situation, we obtain
\[ Z_{n+1} = \varepsilon + \xi_{n+1}Z_n, \]  
(3.11)
with
\[ \xi_n = 1 - (D + 1)\varepsilon + i\zeta_n, \]  
(3.12)
and where \( \zeta_n \) is a Gaussian random variable, such that \( \langle \zeta_n \rangle = 0 \) and \( \langle \zeta_n^2 \rangle = 2D\varepsilon \).

Figure 2 shows a typical trajectory \( \{Z(t)\} \) starting from the origin, of length \( t = 25 \), for \( D = 1 \), i.e. the case of the binary annihilation process (1.7). The trajectory is generated by means of scheme (3.11), with \( \varepsilon = 10^{-3} \). The main advantage of using the complex coordinate \( Z(t) \) (instead of the original field variable \( \varphi(t) \) [7, 8]) is that unbounded excursions are avoided, as we have \( |Z(t)| \leq 1 \). The rest of this section is devoted to the distribution of the functional \( Z \), i.e. equivalently, of the generic point of such a trajectory in the long-time regime.

3.2. Kesten approach

The random variable \( Z \) can be alternatively described as a complex Kesten variable. Along the lines of section 2.2, it can indeed be shown that \( Z \) obeys the identity (in the \( \varepsilon \rightarrow 0 \) limit)
\[ Z \equiv \varepsilon + \xi Z', \]  
(3.13)
where \( Z' \) is a copy of \( Z \), whereas
\[ \xi = e^{-\varepsilon i\zeta}. \]  
(3.14)
Figure 2. A typical trajectory \{Z(t)\} of length \(t = 25\) in the unit disk for \(D = 1\), i.e. the case of the binary annihilation process (1.7).

and \(\zeta\) is Gaussian, such that \(\langle \zeta \rangle = 0\) and \(\langle \zeta^2 \rangle = 2D\varepsilon\). In other words, \(Z\) represents (in the \(\varepsilon \to 0\) limit) the stationary solution of the random recursion

\[ Z_{n+1} = \varepsilon + \xi_{n+1}Z_n, \quad (3.15) \]

with

\[ \xi_n = e^{-\varepsilon+i\zeta_n}. \quad (3.16) \]

The variable \(Z\) therefore identifies (up to an unimportant global factor \(\varepsilon\), and in the \(\varepsilon \to 0\) limit) with the Kesten variable generated by the complex input variables \(\xi_n\).

It is worth noting the close resemblance between the random recursions (3.11) and (3.15), respectively, corresponding to the Langevin and Kesten approaches. Both discrete schemes become equivalent to leading order in the \(\varepsilon \to 0\) limit. The Langevin scheme involving \(\hat{\xi}_n\) appears in this regime as a suitably linearized form of the non-linear Kesten scheme involving \(\xi_n\). The Langevin scheme is more suitable for extensive numerical simulations, as it does not involve the complex exponential function.

3.3. Moments

In the case of the real exponential functional \(X\), the explicit form (2.15) of the distribution \(f_X\) could be derived either from the Langevin approach, as the stationary solution (2.8) of the Fokker–Planck equation, or from the Kesten approach, by means of identity (2.24) in the \(\varepsilon \to 0\) limit. In this case of the imaginary exponential functional \(Z\), neither route leads to an explicit expression for the density \(f(x, y)\).

The Kesten approach however directly yields an interesting piece of information, in the form of a recursion relation for the moments \(\langle Z^k \hat{Z}^l \rangle\). We recall that the distribution of a real random variable \(X\) is entirely characterized by the single array of moments \(\langle X^k \rangle (k = 0, 1, \ldots)\) (leaving aside questions related to convergence), whereas a double array of moments \(\langle Z^k \hat{Z}^l \rangle\)
(or, equivalently, \( (X^kY^l) \)) is needed to characterize the distribution of a complex random variable \( Z = X + iY \).

The analysis of the moments follows the line of thought of section 2.2. For \( k = 0, 1, \ldots \), integer, we expand identity (3.13) as
\[
Z^k = \xi^k(Z')^k + k\xi^{k-1}(Z')^{k-1}e + \cdots. \tag{3.17}
\]
We thus obtain an expansion similar to (2.28), i.e.,
\[
\langle Z^k \bar{Z}^l \rangle = \langle \xi^k \bar{\xi}^l \rangle \langle Z^k \bar{Z}^l \rangle + (k\langle \xi^{k-1} \bar{\xi}^l \rangle \langle Z^k \bar{Z}^l \rangle + l\langle \xi^k \bar{\xi}^{l-1} \rangle \langle Z^k \bar{Z}^l \rangle)\varepsilon + \cdots. \tag{3.18}
\]
Furthermore, we have
\[
\langle \xi^k \bar{\xi}^l \rangle = e^{-(k+l)\varepsilon} \langle e^{i(k-l)\xi} \rangle = e^{-i(k+l+1)\varepsilon} = 1 - (k + l + (k - l)^2 D)\varepsilon + \cdots. \tag{3.19}
\]
Inserting the latter expansion into (3.18), we are left (in the \( \varepsilon \to 0 \) limit) with the recursion
\[
(k + l + (k - l)^2 D)\langle Z^k \bar{Z}^l \rangle = k\langle Z^k \bar{Z}^l \rangle + l\langle Z^k \bar{Z}^l \rangle. \tag{3.20}
\]
Formula (3.20) is one of the key results of this work. It allows one to evaluate the moments \( \langle Z^k \bar{Z}^l \rangle \) in a recursive way. All of them are rational functions of the diffusion constant \( D \). They are regular over the whole range of positive values of \( D \). Owing to the symmetry between \( Z \) and \( \bar{Z} \), we have \( \langle Z^k \bar{Z}^l \rangle = \langle Z^l \bar{Z}^k \rangle \), so that we can restrict ourselves to \( k \geq l \). The expressions of the first moments (for \( k + l \leq 6 \)) are listed in table 1.

The moments \( \langle Z^k \rangle \) can be evaluated in closed form. Equation (3.20) indeed boils down for \( l = 0 \) to the recursion
\[
(kD + 1)\langle Z^k \rangle = \langle Z^{k-1} \rangle, \tag{3.21}
\]
which can be viewed as the analytical continuation of the functional equation (2.30). We thus obtain
\[
\langle Z^k \rangle = \frac{1}{\prod_{j=1}^k (jD + 1)} = \frac{\Gamma(1 + 1/D)}{D^k \Gamma(k + 1 + 1/D)}. \tag{3.22}
\]
For \( D = 1 \), the above expression simplifies to
\[
\langle Z^k \rangle \bigg|_{D=1} = \frac{1}{(k + 1)!}. \tag{3.23}
\]
This result also holds for negative \( k \). Setting \( k = -p \), with \( p \geq 2 \) integer, we obtain \( \langle Z^{-p} \rangle \bigg|_{D=1} = 0 \). We thus recover a result mentioned in section 1, concerning the null factorial moments for the number of particles in the steady state for the binary annihilation process (1.7).

The joint moments \( \langle X^kY^l \rangle \) of the real variables \( X \) and \( Y \) can be obtained as linear combinations of the moments \( \langle Z^k \bar{Z}^l \rangle \), using \( X = (Z + \bar{Z})/2 \) and \( Y = i(Z - \bar{Z})/2 \). The conjugation symmetry ensures that \( f(x, y) = f(x, -y) \), so that \( \langle X^kY^l \rangle = 0 \) whenever \( l \) is odd. The expressions of the first non-vanishing moments \( \langle X^kY^l \rangle \) (for \( l \) even and \( k + l \leq 6 \)) are listed in table 2.

### 3.4. Fokker–Planck equation

The next natural step in our analysis consists in writing down the Fokker–Planck equation obeyed by the distribution of the variable \( Z \) or, equivalently, by the joint distribution of the real variables \( X \) and \( Y \). We shall successively do so using various representations.
Table 1. Expressions of the moments $\langle Z^k \bar{Z}^l \rangle$ for $k + l \leq 6$ as a function of the diffusion constant $D$.

| Moment | Expression |
|--------|------------|
| $\langle Z \rangle$ | $\frac{1}{D+1}$ |
| $\langle Z^2 \rangle$ | $\frac{1}{(D+1)(2D+1)}$ |
| $\langle Z \bar{Z} \rangle$ | $\frac{1}{D+1}$ |
| $\langle Z^3 \rangle$ | $\frac{1}{(D+1)(2D+1)(3D+1)}$ |
| $\langle Z^2 \bar{Z} \rangle$ | $\frac{1}{(D+1)(D+3)(2D+1)}$ |
| $\langle Z^4 \rangle$ | $\frac{1}{(D+1)(2D+1)(3D+1)(4D+1)\frac{6D^2 + 10D + 3}{36D^2 + 93D^2 + 68D + 15}}$ |
| $\langle Z^3 \bar{Z} \rangle$ | $\frac{1}{(D+1)(2D+1)(3D+1)(4D+1)(5D+1)\frac{25D^2 + 16D + 3}{288D^2 + 544D^2 + 279D + 45}}$ |
| $\langle Z^2 \bar{Z}^2 \rangle$ | $\frac{1}{(D+1)(2D+1)(3D+1)(4D+1)(5D+1)\frac{36D^2 + 93D^2 + 68D + 15}{36D^2 + 93D^2 + 68D + 15}}$ |

We start by considering the two-variable characteristic function

$$\Phi(p, q) = \langle e^{pZ+q\bar{Z}} \rangle = \sum_{k,l=0}^{\infty} \langle Z^k \bar{Z}^l \rangle \frac{p^k q^l}{k!l!}. \quad (3.24)$$

The recursion (3.20) is equivalent to the partial differential equation

$$\left( D \left( p \frac{\partial}{\partial p} - q \frac{\partial}{\partial q} \right)^2 + p \frac{\partial}{\partial p} + q \frac{\partial}{\partial q} - p - q \right) \Phi(p, q) = 0, \quad (3.25)$$

with $\Phi(0, 0) = 1$. For $q = 0$, the above equation boils down to the solvable ordinary differential equation

$$Dp\phi''(p) + (D + 1)\phi'(p) - \phi = 0 \quad (3.26)$$

for the characteristic function $\phi(p) = \langle e^{pZ} \rangle$, where accents denote differentiations. We thus obtain

$$\phi(p) = \Gamma(1 + 1/D) \left( \frac{p}{D} \right)^{-1/(2D)} I_{1/2D} \left( 2\sqrt{\frac{p}{D}} \right), \quad (3.27)$$
where $I_{1/6}$ is the modified Bessel function. This result agrees with expression (3.22) of the moments (Z $^k$).

The characteristic function $\Psi(u, v) = \langle e^{iX+Y}\rangle$ is equal to $\Phi(p, q)$, with $p = (u - iv)/2$ and $q = (u + iv)/2$. It therefore obeys the partial differential equation

$$\left(D \left( u \frac{\partial}{\partial v} - v \frac{\partial}{\partial u} \right)^2 - u \frac{\partial}{\partial u} - v \frac{\partial}{\partial v} + u \right) \Psi(u, v) = 0,$$

with $\Psi(0, 0) = 1$. The latter equation is equivalent to the following equation for the joint density of $X, Y$:

$$\left(D \left( y \frac{\partial}{\partial x} - x \frac{\partial}{\partial y} \right)^2 + (x - 1) \frac{\partial}{\partial x} + y \frac{\partial}{\partial y} + 2 \right) f(x, y) = 0.$$

The above equation can be recovered by noticing that $f(x, y)$ is the stationary solution of the two-dimensional (forward) Fokker–Planck equation associated with the two coupled equations. 

| Moment | Expression |
|--------|------------|
| $\langle X \rangle$ | $\frac{1}{D + 1}$ |
| $\langle X^2 \rangle$ | $\frac{1}{2D + 1}$ |
| $\langle Y \rangle$ | $\frac{(D + 1)(2D + 1)}{9D^2 + 10D + 3}$ |
| $\langle X^3 \rangle$ | $\frac{(D + 1)(D + 3)(2D + 1)(3D + 1)}{3D}$ |
| $\langle XY \rangle$ | $\frac{(D + 3)(2D + 1)(3D + 1)}{18D^4 + 60D^3 + 58D^2 + 22D + 3}$ |
| $\langle X^2 Y \rangle$ | $\frac{(D + 1)(D + 3)(2D + 1)(3D + 1)(4D + 1)}{3D^2(6D^2 + 8D + 3)}$ |
| $\langle Y^4 \rangle$ | $\frac{(D + 1)^2(D + 5)(2D + 1)(3D + 1)(4D + 1)5D + 1)}{450D^4 + 940D^3 + 606D^2 + 158D + 15}$ |
| $\langle X^3 Y \rangle$ | $\frac{(D + 1)^2(D + 5)(2D + 1)(3D + 1)(4D + 1)(5D + 1)}{90D^4 + 268D^3 + 249D^2 + 98D + 15}$ |
| $\langle XY^4 \rangle$ | $\frac{(D + 1)^2(D + 5)(2D + 1)(3D + 1)(4D + 1)(5D + 1)}{3D^2(30D^3 + 86D^2 + 63D + 15)}$ |
| $\langle X^4 \rangle$ | $\frac{29D^2 + 1674D^3 + 3594D^3 + 36182D^4 + 19056D^4 + 5383D^2 + 774D + 45}{3D(180D^2 + 756D + 1044D^3 + 584D^2 + 148D + 15)}$ |
| $\langle X^2 Y^4 \rangle$ | $\frac{3D^2(180D^2 + 576D^3 + 638D^2 + 279D + 45)}{3D^2(180D^2 + 576D^3 + 638D^2 + 279D + 45)}$ |
| $\langle Y^6 \rangle$ | $\frac{(D + 1)^2(D + 5)(2D + 1)(3D + 1)(4D + 1)(5D + 1)}{15D^2(180D^2 + 576D^3 + 638D^2 + 279D + 45)}$ |
Langevin–Stratonovich equations
\[ \frac{dX(t)}{dt} = 1 - X(t) - gY(t)\eta(t), \]
\[ \frac{dY(t)}{dt} = -Y(t) + gX(t)\eta(t). \] (3.30)

A single real noise \( \eta(t) \) enters both equations (3.30), and so the above system of Langevin equations is degenerate. As a consequence, the second-order differential operators which appear in either (3.25), (3.28), or (3.29) are parabolic. This phenomenon is best exhibited in yet another form, using the polar coordinates \( Z = R e^{i\theta} \). The system (3.30) then becomes
\[ \frac{dR(t)}{dt} = -R(t) + \cos\theta(t), \]
\[ \frac{d\theta(t)}{dt} = -\sin\theta(t)R(t) + g\eta(t), \] (3.31)
showing explicitly that the noise only acts in the angular direction. This property was already noticed in [7, 8] in the context of binary annihilation (see (1.7)), i.e. for \( D = 1 \).

Coming back to the stationary state, still denoting \( Z = R e^{i\theta} \), and setting
\[ p = \mu \frac{1}{2} e^{-iv}, \]
\[ q = \mu \frac{1}{2} e^{iv}, \] (3.32)
the characteristic function \( \Phi(p, q) \) becomes the generating function
\[ \Omega(\mu, v) = \langle e^{\mu X_v} \rangle, \] (3.33)
where
\[ X_v = \frac{1}{2}(Z e^{-iv} + \bar{Z} e^{iv}) = R \cos(\theta - v) = X \cos v + Y \sin v \] (3.34)
is the projection of \( Z \) onto an axis at an angle \( v \). The function \( \Omega(\mu, v) \) obeys
\[ \left( D \frac{d^2}{dv^2} - \mu \frac{d}{d\mu} + \mu \cos v \right) \Omega(\mu, v) = 0, \] (3.35)
with \( \Omega(0, v) = 1 \) for all \( v \).

The partial differential equation (3.35) is another central result of this work. It will be used in the following to investigate the density of \( Z \) in various limits. We shall discuss the weak-disorder regime \( (D \to 0) \) in section 3.5, the strong-disorder limit \( (D \to \infty) \) in sections 3.6 and 3.9, and the asymptotic behavior of the density of \( Z \) near the unit circle in section 3.8. In spite of its relatively simple form, we have not been able to solve the parabolic partial differential equation (3.35) in full generality. We nevertheless find it worth sketching below two the approaches we have tried.

3.4.1. Fourier-series expansion. The generating function \( \Omega(\mu, v) \) is \( 2\pi \)-periodic in the angular variable \( v \). It is henceforth natural to expand it as a Fourier series:
\[ \Omega(\mu, v) = \sum_{n=-\infty}^{\infty} \omega_n(\mu) e^{in\theta}. \] (3.36)
The Fourier coefficients have the symmetry \( \omega_{-n}(\mu) = \bar{\omega}_n(\mu) \). They obey the coupled differential equations
\[ \left( Dn^2 + \mu \frac{d}{d\mu} \right) \omega_n(\mu) = \frac{\mu}{2} \left( \omega_{n+1}(\mu) + \omega_{n-1}(\mu) \right), \] (3.37)
with \( \omega_n(0) = \delta_{n,0} \). For \( D = 0 \), (3.37) coincides with the differential recursion relation obeyed by the modified Bessel functions \( I_n(\mu) \). Looking for asymptotics similar to those of Bessel
functions, we obtain the following behavior of the amplitudes \( \omega_n(\mu) \) at small and large values of \( \mu \) for all \( n \geq 0 \):

\[
\omega_n(\mu) \approx \frac{\Gamma(1 + 1/D)}{n!\Gamma(n + 1 + 1/D)} \left( \frac{\mu}{2D} \right)^n \quad (\mu \ll 1),
\]

\[
(3.38)
\]

\[
\omega_n(\mu) \sim e^{\mu} \quad (\mu \gg 1).
\]

A more advanced analysis of the large-\( \mu \) regime will be performed in section 3.8.

3.4.2. Expansion in trigonometric polynomials. The generating function \( \Omega(\mu, \nu) \) can alternatively be expanded as a power series in \( \mu \) of the form

\[
\Omega(\mu, \nu) = \sum_{k=0}^{\infty} \frac{\mu^k}{k!} P_k(\cos \nu).
\]

Equation (3.33) yields

\[
P_k(\cos \nu) = \langle X_k \rangle = \langle (X \cos \nu + Y \sin \nu)^k \rangle = \frac{1}{2^k} \langle (Z e^{-i\nu} + \bar{Z} e^{i\nu})^k \rangle.
\]

The \( P_k \) are polynomials of degree \( k \) in \( x = \cos \nu \). They satisfy the differential recursion relation

\[
D((x^2-1)P_k'' + x P_k') + kP_k = kx P_{k-1},
\]

with \( P_0(\nu) = 1 \). Alternatively, as a consequence of (3.40), we have

\[
P_k(\cos \nu) = \frac{1}{2^k} \sum_{j=0}^{k} \binom{k}{j} |Z|^{k-j} |\bar{Z}|^j \cos((2j-k)\nu).
\]

Both (3.41) and (3.42) together with (3.20) or table 1 consistently yield

\[
P_1(x) = \frac{x}{D+1}, \quad P_2(x) = \frac{x^2 + D}{(D+1)(2D+1)}, \quad P_3(x) = \frac{(D+3)x^3 + 9D(D+1)x}{(D+1)(D+3)(2D+1)(3D+1)},
\]

and so on.

The polynomials \( P_k(x) \) do not seem to belong to any of the known standard families of polynomials. In particular, they do not form a set of orthogonal polynomials. This can easily be checked using the fact that orthogonal polynomials must obey a three-term linear recursion [23].

An asymptotic analysis of the large-order behavior of the polynomials \( P_k(x) \) will also be performed in section 3.8.

3.5. Weak-disorder regime (\( D \to 0 \))

In this regime, the complex exponential involved in (1.5) or (3.1) or, equivalently, the trigonometric functions involved in (3.2) can be expanded in powers of \( g \equiv \sqrt{2D} \). We thus obtain

\[
Z = 1 + i\sqrt{2D}U - DV + \cdots,
\]

i.e.

\[
X = 1 - DV + \cdots, \quad Y = \sqrt{2D}U + \cdots.
\]

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where $U$ and $V$ are the following random integrals:

\[
U = \sqrt{2} \int_0^\infty e^{-t} B(t) \, dt, \quad V = \int_0^\infty e^{-t} B(t)^2 \, dt.
\]  

(3.46)

The variable $U$ is Gaussian, and therefore entirely characterized by its second moment $\langle U^2 \rangle = 1$. The random variables $U$ and $V$ however have a non-trivial joint distribution supported by the parabolic domain $V \geq U^2/2$. As $U$ and $V$ are respectively defined as a linear and a quadratic functional of Brownian motion, the characteristic function of their joint distribution can be evaluated either by means of a Gaussian functional integral [24]. An alternative route is provided by the probabilistic approach put forward in [25] and [26, exercise 1.34, p 453], leading to the expression of the joint distribution of arbitrary linear and quadratic functionals of the Brownian motion. Within this framework we prefer to use a more direct approach, based on the solution of the partial differential equation (3.35) in the relevant regime, i.e. $D \ll 1$ and $\nu \ll 1$. We are thus led to consider a simplified version of (3.35), where the cosine function has been expanded to second order:

\[
\left( D \frac{\partial^2}{\partial \nu^2} - \mu \frac{\partial}{\partial \mu} + \mu \left( 1 - \frac{\nu^2}{2} \right) \right) \Omega(\mu, \nu) = 0.
\]  

(3.47)

The latter equation can be solved for arbitrary $D$ if, inspired by the property that $U$ is Gaussian, we make a Gaussian ansatz of the form

\[
\Omega(\mu, \nu) = a(\mu) e^{-b(\mu) \nu^2}.
\]  

(3.48)

We are left with two ordinary differential equations for $a(\mu)$ and $b(\mu)$, namely

\[
b'(\mu) = \frac{1}{2} - \frac{4D}{\mu} b(\mu)^2, \quad a'(\mu) = \left( 1 - \frac{2D}{\mu} b(\mu) \right) a(\mu),
\]  

(3.49)

with $a(0) = 1$ and $b(0) = 0$. The equation for $b(\mu)$ is a Riccati equation, and is therefore solvable. Indeed, setting

\[
b(\mu) = -\frac{1}{8D} + \frac{\mu}{4D} \frac{\psi'(\mu)}{\psi(\mu)},
\]  

(3.50)

we obtain for $\psi(\mu)$ the second-order linear differential equation

\[
\psi''(\mu) = \left( \frac{2D}{\mu} - \frac{1}{4\mu^2} \right) \psi(\mu),
\]  

(3.51)

whose regular solution is

\[
\psi(\mu) = \sqrt{\mu} I_0(2\sqrt{2D\mu}).
\]  

(3.52)

Skipping details, we obtain after some algebra

\[
a(\mu) = \frac{e^{\mu}}{\sqrt{I_0(2\sqrt{2D\mu})}}, \quad b(\mu) = \frac{1}{4D} \sqrt{2D\mu} \frac{I_1(2\sqrt{2D\mu})}{I_0(2\sqrt{2D\mu})},
\]  

(3.53)

where $I_0$ and $I_1$ are modified Bessel functions. In the following we use various properties of Bessel functions (differential equations, series and asymptotic expansions, location of the zeros) presented e.g. in [27].

Now, inserting (3.45) into (3.33) and (3.34), and expanding the result for $D$ and $\nu$ small, we obtain

\[
\Omega(\mu, \nu) = \left\{ \exp \left( \mu + \sqrt{D\mu} \nu U - D\mu V - \frac{1}{2} \mu \nu^2 + \ldots \right) \right\}.
\]  

(3.54)

In the limit where $\mu$ is large, at fixed values of the combinations $\rho = -\sqrt{D\mu} \nu$ and $\sigma = D\mu$, the leading term $\mu$ of the above expansion is divergent, whereas all the other terms remain
finite, and the higher-order terms which are not written down vanish. As a consequence, a direct identification with (3.48), (3.53) yields an explicit expression for the characteristic function

\[ F(\rho,\sigma) = \langle e^{-\rho U - \sigma V} \rangle = \frac{1}{\sqrt{I_0(2\sqrt{2\sigma})}} \exp \left( \frac{I_2(2\sqrt{2\sigma})}{2\sigma I_0(2\sqrt{2\sigma})} \rho^2 \right). \]  

(3.55)

This expression is another key result of this work. It contains the full joint distribution of the rescaled variables \( U \) and \( V \). Table 3 gives the rational values of the first few moments \( \langle U^k V^l \rangle \) (for \( k \) even and \( k, l \leq 4 \)). The positive variables \( U^2 \) and \( V \) appear as strongly positively correlated.

The Gaussian marginal distribution of \( U \) is recovered as \( F(\rho,0) = \langle e^{-\rho U} \rangle = e^{\rho^2/2} \). The marginal distribution of the quadratic functional \( V \) is given by the inverse Laplace transform of \( F(0,\sigma) \), i.e.

\[ f_V(v) = \int d\sigma \frac{e^{\sigma v}}{2\pi i \sqrt{I_0(2\sqrt{2\sigma})}}. \]  

(3.56)

The behavior of this density at small \( v \) is given by a saddle point at a large positive value \( \sigma \approx 1/(2v^2) \) of the integration variable. We thus obtain the exponential fall-off

\[ f_V(v) \sim e^{-1/(2v^2)} \quad (v \to 0). \]  

(3.57)

Its behavior at large \( v \) is determined by the rightmost singularity of the integrand in (3.56), which is a branch-cut singularity at \( \sigma = -\sigma_1 \), with \( \sigma_1 = j_1^2/8 \approx 0.722897 \), where \( j_1 \approx 2.404825 \) is the first zero of the Bessel function \( J_0 \). We thus obtain the exponential fall-off

\[ f_V(v) \sim e^{-\sigma_1 v} \quad (v \to \infty). \]  

(3.58)

Figure 3 shows a plot of the density \( f_V(v) \), obtained by means of a numerical integration of formula (3.56), along the imaginary \( \sigma \) axis.

### 3.6. Strong-disorder regime (\( D \to \infty \))

In this regime, the integrand in expression (1.5) or (3.1) involves a large and hence rapidly varying phase. The variable \( Z \) can therefore be expected to behave as a complex Gaussian variable with zero mean and a small width. Still denoting \( Z = R e^{i\theta} \), we have indeed \( \langle R^2 \rangle = \langle ZZ \rangle \approx 1/D \) (see table 1).

---

**Table 3. Values of the moments \( \langle U^k V^l \rangle \) for \( k \) even and \( k, l \leq 4 \).**

| Moment | Value | Moment | Value | Moment | Value |
|--------|-------|--------|-------|--------|-------|
| \( \langle U^2 \rangle \) | 1 | \( \langle U^4 \rangle \) | 3 |
| \( \langle V \rangle \) | 1 | \( \langle U^2 V \rangle \) | \( \frac{7}{3} \) | \( \langle U^4 V \rangle \) | 11 |
| \( \langle V^2 \rangle \) | 2 | \( \langle U^2 V^2 \rangle \) | \( \frac{25}{3} \) | \( \langle U^4 V^2 \rangle \) | \( \frac{164}{3} \) |
| \( \langle V^4 \rangle \) | \( \frac{20}{3} \) | \( \langle U^2 V^4 \rangle \) | \( \frac{613}{15} \) | \( \langle U^4 V^4 \rangle \) | \( \frac{1726}{5} \) |
| \( \langle V^4 \rangle \) | \( \frac{95}{3} \) | \( \langle U^2 V^4 \rangle \) | \( \frac{2305}{9} \) | \( \langle U^4 V^4 \rangle \) | \( \frac{39751}{15} \) |
Figure 3. Plot of the probability density \( f_V \) of the scaling variable \( V \).

The convergence of \( Z \) to a complex Gaussian variable can be studied by means of the moments investigated in section 3.3. The form of recursion (3.20) implies that the moments behave as \( \langle Z^k \bar{Z}^l \rangle \sim 1/D^{\max(k,l)} \) at large \( D \). As a consequence, the diagonal moments \( \langle R^{2k} \rangle = \langle (Z \bar{Z})^k \rangle \sim 1/D^k \) are the leading ones at large \( D \), whereas the non-diagonal ones \( \langle Z^n (Z \bar{Z})^k \rangle \sim 1/D^{k+n} \) \((n = 1, 2, \ldots)\) are more and more subleading. More precisely, setting

\[
\langle Z^n (Z \bar{Z})^k \rangle = a^{(n)}_k D^{k+n},
\]

recursion (3.20) is equivalent to

\[
d_k^{(0)} = d_{k-1}^{(1)},
\]

\[
(Dn^2 + 2k + n)a_k^{(n)} = D(k + n)a_{k-1}^{(n-1)} + k a_{k-1}^{(n+1)} \quad (n \geq 1).
\]

To leading order as \( D \to \infty \), the latter equation boils down to

\[
n^2 a_k^{(n)} = (k + n)a_{k-1}^{(n-1)},
\]

so that the amplitudes have the asymptotic values

\[
a_k^{(n)} = \frac{(k + n)!}{(n!)^2}.
\]

We thus obtain in particular

\[
\langle (Z \bar{Z})^k \rangle \approx \frac{k!}{D^k}.
\]

We read off from this result that the positive variable \( S = R^2 = Z \bar{Z} = X^2 + Y^2 \) is asymptotically exponentially distributed at large \( D \), with density \( f_S(s) \approx D e^{-Ds} \). In other words, \( Z \) is asymptotically a complex Gaussian variable, with the isotropic density

\[
f(x, y) \approx \frac{D}{\pi} e^{-Dr^2}.
\]

The analysis of recursion (3.60) can be pursued to derive a systematic \( 1/D \)-expansion beyond the asymptotic values (3.62). For the diagonal moments, we thus obtain after some algebra

\[
\langle (Z \bar{Z})^k \rangle \approx \frac{k!}{D^k} \left( 1 - \frac{k(7k + 1)}{8D} + \cdots \right).
\]
3.7. Qualitative features

In this section, we pause in our analytical investigations and turn to a qualitative discussion of various features of the distribution of the variable $Z$, supported by numerical results.

The global picture is provided by figure 4, showing color level plots of the density $f(x,y)$ of the variable $Z = X + iY$ in the unit disk for several values of $D$. The data presented here and throughout the following have been obtained by numerically generating very long time series of the discrete process (3.11), with typically $\varepsilon = 10^{-3}$ and $n = 10^{10}$. First of all, as the distribution is symmetric w.r.t. the $X$-axis, its unique maximum is on the real axis. Figure 5 shows a plot of the observed location $Z_{\text{max}}$ of this maximum (black curve) and of the mean $\langle Z \rangle$ (red line, see (3.6)), against $D/(D + 1)$. Both quantities qualitatively follow the same pattern, decreasing from 1 in the $D \to 0$ limit to 0 in the $D \to \infty$ limit.

The full shape of the distribution is observed to exhibit a smooth dependence on the diffusion constant $D$, continuously interpolating between the weak-disorder and strong-disorder regimes, respectively studied in sections 3.5 and 3.6.

- For small $D$, the distribution of $Z$ is concentrated in a very elongated domain which is stretched along the unit circle and fits its curvature. This very anisotropic distribution is best illustrated on the first panel of figure 4, corresponding to $D = 1/4$. The picture corroborates the fact that $Y$ has a variance $\langle Y^2 \rangle \approx D$, and therefore a vertical extent
scaling as \( \sqrt{D} \), whereas the horizontal extent of \( 1 - X \approx DV \) is much smaller. The curvature reflects the property that \( U^2 \) and \( V \) are strongly positively correlated. 

- For large \( D \), \( Z \) is predicted to be asymptotically a complex Gaussian variable such that \( \langle R^2 \rangle \approx 1/D \). Its distribution is therefore essentially concentrated in a small and nearly circular domain near the origin. This nearly isotropic distribution is illustrated on the last panel of figure 4, corresponding to \( D = 8 \).

The expressions of the variances of \( X \) and \( Y \),
\[
\text{var } X = \frac{D^2}{(D + 1)^2(2D + 1)}, \quad \text{var } Y = \frac{D}{(D + 1)(2D + 1)},
\]
are readily derived from table 1. Both variances go to zero at both endpoints (\( D \rightarrow 0 \) and \( D \rightarrow \infty \)). For small \( D \), \( \text{var } X \approx D^2 \) vanishes much faster than \( \text{var } Y \approx D \), reflecting the very anisotropic nature of the distribution. In contrast, at large \( D \) both variances are asymptotically equal to each other, as \( \text{var } X \approx \text{var } Y \approx 1/(2D) \), reflecting the nearly isotropic nature of the distribution. The variances are maximal for intermediate values of the diffusion constant. The maximum \( (\text{var } X)_{\text{max}} = (5\sqrt{5} - 11)/2 \approx 0.090169 \) of \( X \) is reached for \( D = (1 + \sqrt{5})/2 \approx 1.618033 \), whereas \( (\text{var } Y)_{\text{max}} = 3 - 2\sqrt{2} \approx 0.171572 \) is reached for \( D = 1/\sqrt{2} \approx 0.707106 \). It is worth noting that \( \text{var } Y \) is invariant if \( D \) is changed to its dual \( \tilde{D} \), such that
\[
D \tilde{D} = \frac{1}{2}.
\]
The maximum \( (\text{var } Y)_{\text{max}} \) is consistently reached at the fixed point \( D = 1/\sqrt{2} \) of this duality transformation. Figure 6 shows a plot of both variances against \( D/(D + 1) \).

The features discussed so far are corroborated by the shape of the marginal distributions \( f_X \) and \( f_Y \) of the real variables \( X \) and \( Y \). These distributions, obtained as numerical histograms, are shown in figure 7 for several values of \( D \), chosen so as to form dual pairs (see (3.67)). The dependence of the density of \( X \) on the diffusion constant is observed to evolve smoothly from a narrow distribution around \( X = 1 \) at small \( D \), whose limiting shape is given by the density \( f_V \) of the rescaled variable \( V \), to a narrow Gaussian near \( X = 0 \) at large \( D \), whose variance scales as \( \text{var } X \approx 1/(2D) \). The distribution of \( Y \) behaves differently. It indeed becomes a narrow Gaussian both at weak and at strong disorder, whose variance respectively falls off...
as \( \text{var} Y \approx D \) \((D \to 0)\) and \( \text{var} Y \approx 1/(2D) \) \((D \to \infty)\). The densities of \( Y \) for pairs of dual values \( D \) and \( D' \) of the diffusion constant are observed to be strikingly close to each other.

### 3.8. Asymptotic behavior near the boundary: an essential singularity

We now resume our analytical investigations and turn to more advanced features, namely the asymptotic analysis of the distribution of \( Z \) in two regimes where it is exponentially small: near the boundary (unit circle) for all values of \( D \) (in this section), and for large \( D \) at any point inside the unit disk (in section 3.9).

For the time being, our goal is to investigate the distribution of \( Z \) near the unit circle. More precisely, we are interested in the behavior of the density \( f_\nu(X_\nu) \) of the variable \( X_\nu \) defined in
as the unit circle is approached from inside \((X_\nu \to 1)\) for a fixed value of the angle \(\nu\). Definition (3.33) implies

\[
 f_\nu(X_\nu) = \int \frac{d\mu}{2\pi i} e^{-\mu X_\nu} \Omega(\mu, \nu).
\]  

We are thus led to study the asymptotic behavior of the characteristic function \(\Omega(\mu, \nu)\) at large \(\mu\) and fixed \(\nu\).

It is worthwhile to first have a look at the simplified situation considered in section 3.5, i.e. \(\nu \ll 1\). In this regime, inserting (3.53) into (3.48), and expanding the result for large \(\mu\), we obtain

\[
 \Omega(\mu, \nu) = \exp \left\{ \mu - \sqrt{2D\mu} \left( 1 + \frac{\nu^2}{4D} \right) + \frac{1}{8} \ln(32\pi^2 D\mu) - \frac{\nu^2}{16D} + \cdots \right\}.
\]  

Inspired by the above expansion, we look for a solution of the full partial differential equation (3.35) which behaves at large \(\mu\) as

\[
 \Omega(\mu, \nu) = \exp(\mu - A(\nu) \sqrt{\mu} + \cdots).
\]  

Inserting this expansion into (3.35), we are left with the ordinary differential equation

\[
 \left( \frac{dA}{d\nu} \right)^2 = 1 - \cos \nu = \frac{2}{D} \sin^2 \frac{\nu}{2}
\]  

for the amplitude \(A(\nu)\). Let us assume that the value \(A(0) = \sqrt{2D}\) at \(\nu = 0\), which can be read off from (3.69), holds for all values of \(D\). This hypothesis will be tested against numerical results in a while. We thus obtain

\[
 A(\nu) = \sqrt{\frac{2D}{D + 2 - 2 \cos \frac{\nu}{2}}}.
\]  

The fall-off of the distribution of \(Z\) near the unit circle can be estimated by inserting (3.70) into (3.68):

\[
 f_\nu(X_\nu) \sim \int \frac{d\mu}{2\pi i} e^{\mu(1-X_\nu)-A(\nu)\sqrt{\mu}}.
\]  

For \(X_\nu \to 1\), this integral is dominated by a saddle point at \(\mu \approx A(\nu)^2/(4(1-X_\nu)^2)\). This yields the estimate

\[
 f_\nu(X_\nu) \sim \exp \left( -\frac{A(\nu)^2}{4(1-X_\nu)} \right),
\]  

i.e.

\[
 f_\nu(X_\nu) \sim \exp \left( -\frac{(D + 2 - 2 \cos (\nu/2))^2}{2D(1 - X_\nu)} \right).
\]  

This result confirms that the distribution of \(Z\) is supported by the whole unit disk and predicts that it vanishes exponentially fast near any point of the unit circle. The behavior of the densities of the variables \(X\) and \(Y\) near their endpoints is given by the following simpler formulas, respectively, corresponding to \(\nu = 0\), \(\nu = \pm \pi\), and \(\nu = \pm \pi/2\):

\[
 f_X(x) \sim \exp \left( -\frac{D}{2(1-x)} \right), \quad (x \to 1),
\]

\[
 f_X(x) \sim \exp \left( -\frac{(D + 2)^2}{2D(1+x)} \right), \quad (x \to -1),
\]

\[
 f_Y(y) \sim \exp \left( -\frac{(D + 2 - \sqrt{2})^2}{2D(1 - |y|)} \right), \quad (y \to \pm 1).
\]
Figure 8. Plot of the amplitude $A(\nu)$ against the reduced angle $\nu/\pi$ for three values of the diffusion constant $D$. Full lines: analytical prediction (3.72). Symbols: result of a fit of the first 1000 polynomials $P_k(\cos \nu)$ (see the text).

The first of these estimates matches the behavior (3.57) of the density of $V$.

The dependence of the amplitude $A(\nu)$ on $D$ is not monotonic. It diverges at both endpoints ($D \to 0$ and $D \to \infty$), and therefore reaches a minimum, $A_{\text{min}}(\nu) = 4\sqrt{1 - \cos(\nu/2)}$ for an angle-dependent value of the diffusion constant $D_{\text{min}}(\nu) = 2 - 2\cos(\nu/2)$. This non-monotonic behavior holds for any angle except along the real axis ($\nu = 0$), where we have $A(0) = \sqrt{2D}$. It is reminiscent of the behavior of var $Y$ discussed in section 3.7.

The essential singularity (3.75) of the distribution of $Z$ near the unit circle is unobservable in numerical data. Nevertheless, it has an observable consequence. It indeed reflects itself in the large-order behavior of the polynomials $P_k(\cos \nu) = \langle X_k^\nu \rangle$ (see (3.40)). We have

$$P_k(\cos \nu) = \int_{-1}^{1} X_k^\nu f_\nu(X_\nu) \, dX_\nu. \quad (3.77)$$

For large $k$, the integral is dominated by either of the endpoints. For an angle $\nu < \pi/2$, the leading endpoint is $X_\nu \to 1$, so that the integral can be estimated as

$$P_k(\cos \nu) \sim \int_{-1}^{1} X_k^\nu \exp\left(-\frac{A(\nu)^2}{4(1 - X_\nu)}\right) \, dX_\nu. \quad (3.78)$$

The integral is dominated by a saddle point at $X_\nu \approx 1 - A(\nu)/(2\sqrt{k})$, yielding

$$P_k(\cos \nu) \sim \exp(-A(\nu)\sqrt{k}). \quad (3.79)$$

We therefore predict a stretched exponential decay of the polynomials $P_k(\cos \nu)$, at any fixed angle $\nu$. Such a behavior can easily be observed numerically. Recursion (3.20) and expansion (3.42) indeed allow a very accurate numerical evaluation of at least the first 1000 polynomials. Figure 8 shows a plot of the amplitude $A(\nu)$ against the reduced angle $\nu/\pi \leq 1/2$ for several values of the diffusion constant $D$. The striking agreement between the prediction (3.72) for the amplitude (full lines) and the result of a quadratic fit of all the first 1000 polynomials as $\ln P_k(\cos \nu) = -A\sqrt{k} + B \ln k + C$ (symbols) shows in particular that the hypothesis made above (3.72) is valid.
3.9. Large deviations at large $D$

We close our investigation by returning to the regime where the diffusion constant $D$ is large. In this situation, it has been shown in section 3.6 that the bulk of the distribution of $Z$ is a narrow isotropic Gaussian, whose density is proportional to $e^{-DR^2}$ (see (3.64)). Furthermore, the estimate (3.75) of the distribution near the unit circle also becomes isotropic at large $D$, and it falls off at $e^{-D/(2(1-R^2))}$. These observations suggest that the distribution follows an exponential large-deviation estimate of the form

$$f(x, y) \approx f(R) \sim e^{-DS(R)}$$

(3.80)

for large $D$, all over the unit disk.

Our goal is to show that such an estimate indeed holds and to derive the large-deviation function $S(R)$. First, if the distribution of $Z$ is nearly isotropic, we have

$$\Omega(\mu, \nu) \approx 2\pi \int_0^\infty Rf(R)I_0(\mu R)\,dR$$

(3.81)

for all values of $\nu$. Then, if the distribution obeys the formula (3.80), the above integral can be evaluated for large $\mu$ (i.e. $\mu$ of the order of $D$) by means of the saddle-point approximation. Setting $\mu = Dx$, we obtain

$$\Omega(\mu, \nu) \sim e^{DS(x)}$$

(3.82)

where the functions $S(R)$ and $\Sigma(x)$ are the Legendre transforms of each other:

$$\Sigma(x) + S(R) = xR, \quad x = S(R), \quad R = \Sigma'(x).$$

(3.83)

Formulas (3.80) and (3.82) are just exponential estimates. In order to proceed, we have to consider the angular dependence of the pre-exponential factor in the latter estimate. Setting

$$\Omega(\mu, \nu) \approx H(\nu) e^{DS(x)}$$

(3.84)

the partial differential equation (3.35) yields the ordinary differential equation

$$H''(\nu) + (x \cos \nu - E)H(\nu) = 0$$

(3.85)

for the amplitude $H(\nu)$ at fixed $x = \mu/D$ with $E = xR = \mu R/D$. Equation (3.85) is known as the Mathieu equation [27]. The requirement that $H(\nu)$ be periodic and positive implies that $E(x)$ is the ground-state eigenvalue of the latter equation. Introducing the Fourier series

$$H(\nu) = \sum_{n=-\infty}^\infty h_n e^{in\nu},$$

(3.86)

the Mathieu equation (3.85) is equivalent to the three-term recursion

$$2(n^2 + E)h_n = x(h_{n+1} + h_{n-1})$$

(3.87)

which is the appropriately rescaled form of recursion (3.37). The ground-state eigenvalue $E(x)$ is characterized by the property that the above recursion has a positive normalizable solution. The function $E(x)$ being known, $\Sigma(x)$ is obtained by integrating the relation $\Sigma'(x) = E(x)/x$ with $\Sigma(0) = 0$, and finally we have $R = E(x)/x$ and $S(R) = E(x) - \Sigma(x)$. We have thus obtained the large-deviation function $S(R)$ in the parametric form. Analytical expansions can be performed for $x \to 0$ and $x \to \infty$, respectively, corresponding to $R \to 0$ and $R \to 1$.

For $x \to 0$, the amplitudes $h_n \sim x^n$ of the harmonics fall off very fast. Taking successive harmonics into account, we can derive the expansions

$$E(x) = \frac{1}{2}x^2 - \frac{7}{12}x^4 + \frac{29}{72}x^6 + \cdots \quad (x \to 0),$$

$$S(R) = R^2 + \frac{7}{8}R^4 + \frac{395}{432}R^6 + \cdots \quad (R \to 0).$$

(3.88)

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For $x \to \infty$, the function $H(\nu)$ becomes peaked around $\nu = 0$. In this regime, (3.85) maps onto a weakly anharmonic quantum-mechanical oscillator. Treating the leading nonlinearity in $\nu^4$ to first order, we obtain

$$E(x) = x - \sqrt{\frac{x}{2}} + \frac{1}{16} + \cdots \quad (x \to \infty),$$

$$S(R) = \frac{1}{2(1 - R)} + \frac{1}{8} \ln(1 - R) + C + \cdots \quad (R \to 1).$$

(3.89)

The leading behavior of $S(R)$ at both endpoints matches the known results recalled at the beginning of this section. The coefficient $1/8$ of the logarithmic term in $S(R)$ near $R = 1$ matches the similar coefficient in (3.69).

The above scheme is easily implemented numerically. This yields e.g. the constant $C \approx -0.603$. Figure 9 shows a plot of the function $S(R)$ thus obtained (left) and of the small difference $\delta S(R) = S(R) - R^2/(1 - R^2)$ (right).

4. Discussion

This work has been devoted to a detailed investigation of a random complex integral $Z$, which provides the most natural example of an imaginary exponential functional of Brownian motion. Both in the warming-up part (section 2) devoted to the real functional $X$ and in the main part (section 3) devoted to $Z$, the main emphasis has been put on the complementarity between the more traditional approach, based on Langevin equations and stochastic calculus, and a more original one, based on discrete random recursions and Kesten variables. Even though neither of these routes leads to the solution of the problem, i.e. to an explicit expression for the probability distribution function of $Z$, we have gathered many partial results concerning various facets of this distribution in various regimes. Let us emphasize in particular the recursion relation (3.20) for the moments $\langle Z^k \bar{Z}^l \rangle$ and the occurrence of an essential singularity of the form (3.75) of the asymptotic behavior of the density near the unit circle.

Apart from its intrinsic interest, the imaginary exponential functional $Z$ has close connections with the realm of reaction-diffusion processes. This fact, already underlined in [8], was our original motivation for pursuing this work. The ubiquity of imaginary noise Langevin equations illustrates the potential pitfalls one faces when building up phenomenological
stochastic equations in out-of-equilibrium situations. In fact, within the Doi–Peliti formalism for reaction-diffusion processes, it can be shown [28] that degenerate Langevin equations for a complex-valued field are the rule rather than the exception. This study will hopefully trigger further interest in such stochastic processes.

To close, it is worth mentioning two recent articles whose subject matters are related to this work. First, explicit examples of invariant measures with complex support generated by random products of complex $2 \times 2$ matrices have been given by Marklof et al [29]. More recently, Dobrinevski et al [30] have considered the motion of a particle in a complex random landscape corresponding to amplitude and phase disorder. The case referred to by those authors as imaginary Brownian disorder has some overlap with this problem and the results match whenever they should do.

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