Interactions of BDNF Val66Met Polymorphism and Menstrual Pain on Brain Complexity

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The irregularity and uncertainty of neurophysiologic signals across different time scales can be regarded as neural complexity, which is related to the adaptability of the nervous system and the information processing between neurons. We recently reported general loss of brain complexity, as measured by multiscale sample entropy (MSE), at pain-related regions in females with primary dysmenorrhea (PDM). However, it is unclear whether this loss of brain complexity is associated with inter-subject genetic variations. Brain-derived neurotrophic factor (BDNF) is a widely expressed neurotrophin in the brain and is crucial to neural plasticity. The BDNF Val66Met single-nucleotide polymorphism (SNP) is associated with mood, stress, and pain conditions. Therefore, we aimed to examine the interactions of BDNF Val66Met polymorphism and long-term menstrual pain experience on brain complexity. We genotyped BDNF Val66Met SNP in 80 PDM females (20 Val/Val, 31 Val/Met, 29 Met/Met) and 76 healthy female controls (25 Val/Val, 36 Val/Met, 15 Met/Met). MSE analysis was applied to neural source activity estimated from resting-state magnetoencephalography (MEG) signals during pain-free state. We found that brain complexity alterations were associated with the interactions of BDNF Val66Met polymorphism and menstrual pain experience. In healthy female controls, Met carriers (Val/Met and Met/Met) demonstrated lower brain complexity than Val/Val homozygotes in extensive brain regions, suggesting a possible protective role of Val/Val homozygosity in brain complexity. However, after experiencing long-term menstrual pain, the complexity differences between different genotypes in healthy controls were greatly diminished in PDM females, especially in the limbic system, including the hippocampus and amygdala. Our results suggest that pain experience preponderantly affects the effect of BDNF Val66Met polymorphism on brain complexity. The results of the present study also highlight the potential utilization of resting-state brain complexity for the development of new therapeutic strategies in patients with chronic pain.

Keywords: BDNF Val66Met polymorphism, primary dysmenorrhea, brain complexity, multiscale entropy, magnetoencephalography, chronic pain
INTRODUCTION

Chronic pain can be considered as “pain that persists for a given length of time,” where the length of time is determined by common medical experience (Merskey and Bogduk, 2002), and have pronounced female predominance (Mogil, 2012). Primary dysmenorrhea (PDM) indicates “pain with menstruation not associated with a well-defined pathology”¹ and is classified as chronic pelvic pain syndrome (IASP, 2011). It has a prevalence of 45–95% among reproductive age females (Berkley, 2013; Iacovides et al., 2015). Most importantly, dysmenorrhea early in life often co-occurs with many chronic pain syndromes later in life, linking its susceptibility to the development of chronic pain conditions (Berkley, 2013) and marking the assaults of central sensitization on dysfunctional pain modulatory system (Nijs et al., 2015).

Long-term PDM serves as a genuine model to study clinical pain with its natural cyclic painful (menstruation) and pain-free (periovulatory) states (Wei et al., 2016a). Structural and functional brain alterations are reported in PDM females (PDMs), including hypotrophic and hypertrophic changes in gray matter volume (Tu et al., 2010, 2013; Liu et al., 2016), white matter microstructural alterations (Liu J. et al., 2017), maladaptive descending pain modulatory system (Wei et al., 2016a), shift of functional connectivity between resting-state networks (Wu et al., 2016; Liu P. et al., 2017), increased theta activity (Lee et al., 2017), and loss of brain complexity (Low et al., 2017) in brain regions related to sensory, affective, and cognitive dimensions of pain. Notably, genetic polymorphisms have been implicated to contribute to inter-subject variations in susceptibility to menstrual pain (Lee et al., 2014; Wei et al., 2016b, 2017), inviting more studies of neuroimaging genetics in PDM.

Brain-derived neurotrophic factor (BDNF) is the most expressed neurotrophin in the brain, especially in the cerebral cortex and hippocampus (Binder and Scharffman, 2004; Tapié-Arancibia et al., 2004). BDNF has pleiotropic roles in the central nervous system, including neurogenesis, neuronal growth, maturation, survival, synaptic plasticity, and microarchitectural integrity (Park and Poo, 2013; Bathina and Das, 2015). It is a driving force behind neural plasticity and protects neuronal cells upon adverse circumstances (Bathina and Das, 2015), such as stress or pain. BDNF is considered as a pain modulator given its participation in activity-dependent synaptic plasticity within the pain modulatory circuitry (Merighi et al., 2008; Haas et al., 2010; Caumo et al., 2016; Generaal et al., 2016).

Human BDNF gene at chromosome 11p14.1 displays a variety of polymorphisms. The replacement of Valine (Val) by Methionine (Met) at codon 66, namely BDNF Val66Met single-nucleotide polymorphism (rs66265), is considered to disrupt normal trafficking of BDNF and consequently reduces activity-dependent secretion of BDNF and BDNF activity in Met carriers (Egan et al., 2003). Animal studies reported that spike-timing-dependent plasticity in the pyramidal neurons of the infralimbic

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2017; Lebedev et al., 2016; Li et al., 2016; Zhou et al., 2016; Jia et al., 2017; Nelson et al., 2017; Chang et al., 2018; Song et al., 2018). It is noteworthy that multiscale entropy (MSE) analysis (Costa et al., 2002, 2005; Yang and Tsai, 2013; Courtiol et al., 2016) calculates a series of sample entropy over multiple time scales, which captures the temporal complexity characteristics of time-series neural signals from microscopic to macroscopic aspects. Recently, MSE analysis has also been applied to brain signals (Heisz and McIntosh, 2013; Yang et al., 2013; Courtiol et al., 2016), pain (Sitges et al., 2010; Valencia et al., 2016; Liu Q. et al., 2017), and PDM studies (Kuo et al., 2017; Low et al., 2017). By applying MSE analysis on resting-state magnetoencephalography (MEG) signals acquired from PDMs during pain-free state, we observed a general loss of regional complexity in PDMs at brain regions related to chronic pain, including the limbic circuitry, default mode network, sensorimotor network, and salience network (Low et al., 2017). Our findings implicated the assaults of long-term menstrual pain on brain complexity and adaptability. However, it is unclear whether this loss of brain complexity in PDMs is associated with genetic variations.

Long-term menstrual pain is a chronic stressor to PDMs that might affect the secretion levels of BDNF and subsequent BDNF functions in an activity-dependent manner. Given the loss of brain complexity in PDMs and the effects of BDNF Val66Met polymorphism on mood and resting-state functional connectivity in PDMs, we aimed to examine the interactions of BDNF Val66Met polymorphism and long-term menstrual pain experience on brain complexity. We hypothesized that there might be genotype-specific complexity differences in healthy controls, and such complexity differences might be affected by long-term menstrual pain especially in pain- and emotion-related brain circuits.

METHODS

Participants

The participants were a subset of the participants from our multimodal imaging genetics (magnetic resonance imaging and magnetoencephalography) and behavioral studies of PDM at Taipei Veterans General Hospital in Taiwan (Lee et al., 2014; Wei et al., 2016a; Wu et al., 2016; Low et al., 2017) who were eligible for neuroimaging studies. Written informed consent form and psychological inventories were approved by the ethics committee of Institutional Review Board of Taipei Veterans General Hospital, Taiwan. Before the study, all participants who were assessed for eligibility signed the written informed consent form. Studies were conducted in accordance with the Declaration of Helsinki.

The inclusion criteria for PDMs were (1) 20 to 30 years old Taiwanese (Asian) females; (2) 27 to 32 days of regular menstrual cycle; (3) right-handedness assessed by the Edinburgh Handedness Inventory; (4) menstrual pain history longer than half year; (5) averaged menstrual pain rating within the last 6 months of experiment was higher than four out of ten using verbal numerical rating scale (0 = no pain, 10 = worst imaginable pain); (6) no pelvic pathologies examined using pelvic ultrasonography and diagnosed as PDM by gynecologist. We excluded volunteers with (1) organic pelvic diseases; (2) pituitary gland pathologies; (3) history of neurological or psychiatric disorders; (4) history of brain surgery or trauma; (5) history or immediate plans for pregnancy or childbirth; (6) history of using medications or supplements of hormonal therapy including oral contraceptives, central-acting medication, or Chinese herbal medicine within the last 6 months of experiment; (7) claustrophobia; (8) contraindications to magnetic resonance imaging. Also, no painkillers were used 24 h before the experiment. Healthy female controls (CONs) had the same inclusion and exclusion criteria except they had no lower abdominal pain during the menstrual period.

Genotyping

Whole blood was collected during the inception stage and stored in 4 mL EDTA tubes at 4°C refrigerator. DNA was extracted using the Puregene kit following the manufacturer’s guidelines (Genta Systems, Inc., Minneapolis, MN, USA). Genotyping was conducted using commercial TaqMan single-nucleotide polymorphism assays (Applied Biosystems, Inc., Foster City, CA, USA). The polymerase chain reaction amplification protocol was as follows: 10 µL; 50°C (2 min), 95°C (10 min), 40 cycles of 92°C (15 s), and 60°C (1 min). Fluorescence measurements were done using the ABI HT7900 (Applied Biosystems, Inc.). Allele calling was performed by the SDS 2.2 software package (Applied Biosystems, Inc.). Two independent technicians blinded to the participants’ personal information assigned the genotypes.

Demographic Data, Pain Experiences, and Psychological Characteristics

Demographic data included age, body mass index (BMI), and handedness. Menstrual features included age at menarche, years of menstruating, and averaged menstrual cycle length. All participants completed the Chinese version of Basic Personality Inventory (BPI; Wu et al., 1999) to assess their personality traits. There are several scale clusters in the BPI, including the personal emotional adjustment scale cluster (depression, anxiety, and hypochondriasis scales), which is of particular interest of this study. The IQOLA SF-36 Taiwan Standard Version 1.0 (SF-36; Tseng et al., 2003) was used to assess long-term physical and mental quality of life. Since emotional perception and pain chronification can be exacerbated by anxiety and depression and both have been linked to PDM, anxious and depressive moods were investigated using Chinese versions of Spielberger State-trait Anxiety Inventory (STAI; Ma et al., 2013), Beck Depression Inventory (BDI-IA; Beck et al., 1979), and Beck Anxiety Inventory (BAI; Lin, 2000). We also studied pain catastrophizing cognitive style, which is the negative appraisal style of pain, using the Chinese version of Pain Catastrophizing Scale (PCS; Yap et al., 2008).

Menstrual pain experiences were evaluated only in PDMs using the Chinese version of McGill Pain Questionnaire (MPQ; Melzack, 1975, 1983) and verbal numerical rating scale (VNRS). MPQ classifies four categories of qualities of pain, including sensory, affective, evaluative, and miscellaneous. MPQ scores are calculated as the sum of the rank values of the words chosen, summing up to a pain rating index (PRI) for each category. The
MPQ present pain index (PPI), based on a 0-to-5 intensity scale (0 = no pain, 1 = mild, 2 = discomforting, 3 = distressing, 4 = horrible, 5 = excruciating), was used as the indicator of menstrual pain intensity. VNRS is a verbal report of menstrual pain intensity rated from 0 to 10 (0 = not at all, 10 = the worst imaginable pain). PDMs recalled different aspects of their overall menstrual pain experiences over the last 6 months, yielding recalled MPQ scores and recalled pain scores.

Data Acquisition

Resting-State MEG Signals Acquisition

Three-minute eye-closed resting-state MEG signals were recorded using a whole-head 306-channel neuromagnetometer (Vectorview, Elekta Neuromag, Helsinki, Finland) comprising 102 triple sensors (two orthogonal planar gradiometers and one magnetometer at each triple-set) at Taipei General Veterans Hospital, Taiwan. Electrooculography (EOG) was recorded using two vertical and two horizontal electrodes to be used for rejection of epochs coinciding with blinks and excessive eye movements with an amplitude cut-off of 600 μV. Locations of three anatomical landmarks (nasion and two bilateral preauricular points) were identified with a three-dimensional digitizer (Isotraf 3S10002, Polhemus Navigation Sciences, Colchester, Vermont USA) to align MEG coordinate system with MRI coordinate system. Four head position indicator (HPI) coils were used to trace the position of subject's head in the MEG system. The online sampling rate was 1,000 Hz, and online bandpass filter was between 0.03 and 330 Hz with a 60 Hz notch filter. Signals exceeding 6,000 μT/cm were rejected. MEG signals recorded from 204 planar gradiometers were further analyzed; MEG signals recorded from 102 magnetometers were excluded from further analyses due to their susceptibility to distant noises. Participants sat comfortably in a magnetically shielded room (Euroshield, Eura, Finland) with heads covered by the helmet and were instructed to relax, eliminate eye movements, and focus only on their breathing (Low et al., 2017).

Structural MRI Images Acquisition

T1-weighted brain images were acquired using a 3 Tesla magnetic resonance imaging (MRI) scanner (Magnetom Trio Tim, Siemens, Erlangen, Germany) at National Yang-Ming University, Taiwan with 12-channel head coil and standard three-dimensional magnetization-prepared rapid gradient-echo (3D MP-RAGE) sequence. The parameters were as follows (Low et al., 2017): TR = 2530 ms, TE = 3.03 ms, TI = 1,100 ms, flip angle = 7°, field-of-view (FOV) = 224 × 256 mm², number of slices = 192, matrix size = 224 × 256, thickness = 1 mm.

Brain Region Parcellation

A total of 90 cortical regions (45 regions in each hemisphere) were defined using the automated anatomical labeling (AAL) template (Tzourio-Mazoyer et al., 2002) with a spatial resolution of 1 × 1 × 1 mm provided in MRICro freeware (Rorden and Brett, 2000). We first normalized the Montreal Neurological Institute (MNI) template to individual’s MRI images using IBASPM (Individual Brain Atlases using Statistical Parametric Mapping; Alemán-Gómez et al., 2006). The AAL template was subsequently transformed into individual space using the estimated deformation field. Hence, parcellation of 90 brain regions in the individual brain was obtained, and voxel-wise source analyses (section Source Analyses) and MSE calculations (section Multiscale Sample Entropy) were carried out in each brain region. The purpose of this procedure was to avoid interpolation of functional activity and thus preserved the precision of MSE statistical analysis in individual space. Finally, eight resting-state networks based on literature were discussed, including the limbic, default mode, salience, sensorimotor, executive control, attention, visual processing, and auditory processing networks (Low et al., 2017).

Source Analyses

MEG signals from each participant were preprocessed before voxel-wise source reconstruction. The signals were segmented into non-overlapping epochs of 8 s (Low et al., 2017). Artifact rejection threshold was set to 2,000 μT/cm, and EOG rejection threshold was set to 250 μV. Any epoch with amplitude larger than these thresholds was excluded, resulting in around 14 eight-second epochs remained. Signal space projected MEG signals were subsequently band-pass filtered within 0.5–90 Hz (Low et al., 2017) and with a notch filter of 55–65 Hz. Zero-mean adjustment was also applied.

After co-registering the coordinate systems between individual’s MRI volume and MEG device, the source activity was estimated with an isotropic resolution of 4 mm. Maximum contrast beamformer (MCB) was used for brain source calculation (Chen Y. S. et al., 2006). To estimate source activity \( y_v(t) \) at each location \( v \), a spatial filter \( w_v \) was applied to the MEG recordings \( m(t) \):

\[
y_v(t) = w_v^T m(t). \tag{1}
\]

For each location, \( w_v \) was obtained by minimizing variance in the output signal \( y_v(t) \) with the unit-gain constraint \( w_v^T I_v = 1 \), where \( I_v \) is the lead field vector. The details of the calculation can be found in our recent report (Kuo et al., 2017). Source activity of all voxels in every brain region was used in the following analysis.

Multiscale Sample Entropy

Multiscale sample entropy (MSE) was proposed to measure the complexity with multiple time scales (Costa et al., 2002) by applying the sample entropy (SE) method (Richman and Moorman, 2000) to different time scales. First, the coarse-grained time series \( z_{vT} = [z_{vT}(1), z_{vT}(2), \ldots, z_{vT}(N/T)] \) of the original time series \( y_v = [y_v(1), y_v(2), \ldots, y_v(N)] \) with \( N \) sample points was obtained for each scale factor \( T \) (\( T \) was set from 1 to 100 in this study; Low et al., 2017):

\[
z_{vT}(k) = \frac{1}{T} \sum_{i=k-1}^{k} y_v(i), 1 \leq k \leq N/T. \tag{2}
\]

Then the SE method was applied to each \( z_{vT} \). In the SE algorithm, a set of vector \( g_i \) with \( m \) elements can be defined as follows (\( m = 2 \) in this study; Low et al., 2017):

\[
g_i = [z_{vT}(i), z_{vT}(i+1)], \text{ where } i = 1, \ldots, N-2. \tag{3}
\]
For each $i$, the absolute difference between $g_i$ and $g_j$ ($1 \leq j \leq N - 2$ and $j \neq i$) was calculated, and $c_r^h$, the number of the difference smaller than $r$, was determined. Here, the value $r$ was set as $0.25 \times$ standard deviation of the signal $g_i(t)$ (Low et al., 2017). Then, the extended vectors with $m+1$ elements were defined as follows:

$$h_i = [z_r^i(i), z_r^i(i + 1), z_r^i(i + 2)],$$

where $i = 1, \ldots, N - 3$, (4)

and $c_r^h$ was determined by the same means. Finally, MSE for time scale $\tau$ can be calculated as:

$$\text{MSE}_\tau = -\ln \frac{c_r^h}{c_r^\tau}. \quad \text{(5)}$$

The MSE value of each voxel was calculated and then averaged within each region. Figure 1 illustrates examples of MEG signals at different representative time scale factors ($\tau$) in one region (the left amygdala) from one representative subject in each group.

### Statistical Analyses

There were two between-group factors in this study: group (PDM and CON) and BDNF Val66Met genotype (Val/Val, Val/Met, and Met/Met). Statistical analyses of genotypes, demographic data, psychological characteristics, and correlation analyses were performed in IBM SPSS Statistics. Statistical analyses of multiscale sample entropy were performed in Matlab.

### BDNF Genotype Distributions and Allele Frequency

The Hardy-Weinberg equilibrium of the BDNF Val66Met genotype distributions and allele frequency were tested using chi-square tests of goodness-of-fit ($p < 0.05$). The associations between BDNF genotype and PDM were tested using chi-square tests of independence ($p < 0.05$) under SPSS binary logistic regression test, in which PDM (group) was treated as the dependent (outcome) variable, genotype as the categorical predictor variable, and Val/Val, Val/Met, or Val carriers as the reference groups.

### Demographic Data, Pain Experiences, and Psychological Characteristics

Descriptive and normality tests (Shapiro-Wilk test, $p < 0.05$, two-tailed) were performed by groups before any inferential statistical tests (Ghasemi and Zahediasl, 2012). As many of the demographic and psychological scores were not normally distributed, we used non-parametric inferential statistical tests. Group difference of each genotype was tested using Mann-Whitney $U$ test ($p < 0.01$, two-tailed). Genotype differences in each group were tested using Kruskal-Wallis $H$ tests ($p < 0.01$, two-tailed) for continuous data and chi-square tests ($p < 0.01$) for categorical data. As there is no corresponding nonparametric two-way ANOVA test, significant group and genotype main effects and group $\times$ genotype interactions were tested using two-way ANOVAs ($p < 0.01$, two-tailed). Post-hoc pairwise comparisons were performed using Dunn-Bonferroni adjusted $p < 0.01$ (two-tailed).

### Multiscale Sample Entropy

For the averaged MSE value in each brain region, the main effects and interactions of group and genotype were tested using two-way ANOVA (Bonferroni-adjusted $p < 0.05$, two-tailed). Due to possible errors induced by the above-mentioned normalization procedure (section Brain Region Parcellation), we excluded the brain regions that contained <10 voxels. We focused on testing the post-hoc planned pairwise comparisons between groups for each genotype and comparisons between genotypes for each group using permutation tests (iterations = 5000, $p < 0.005$, two-tailed). We also tested post-hoc planned comparisons using the more conservative Bonferroni-correction (Bonferroni-adjusted $p < 0.05$, two-tailed).

MSE of the brain regions that showed significant group differences in the same genotype group (Val/Val: PDM vs. CON; Val/Met: PDM vs. CON; Met/Met: PDM vs. CON) were termed as “pain-associated regional MSE.” MSE of the brain regions that showed significant genotype differences (Val/Val vs. Met/Met, Val/Val vs. Met/Met vs. Met/Met) in the same group were termed as “BDNF-associated regional MSE.” To examine the effect sizes of pain-associated regional MSE and BDNF-associated regional MSE, standardized effect sizes were calculated using Cohen’s $d$ (Cohen, 1988) with an online effect size calculator (Wilson and Lipsey, 2001).

### Correlations Between Regional MSE and Psychological Characteristics

Correlations between PDMs’ menstrual pain experiences/psychological characteristics and pain-associated or BDNF-associated regional MSE were examined using Spearman correlation analysis ($p < 0.01$, two-tailed).

### RESULTS

#### BDNF Val66Met Genotype Distributions and Allele Frequency

Genotype distributions of BDNF Val66Met (rs6265) in the CON group, PDM group, and all participants were in Hardy-Weinberg equilibrium ($p > 0.05$; Table S1). The number of participants in each genotype subgroup were as follows (Table 1): Val/Val CONs = 25, Val/Met CONs = 36, Met/Met CONs = 15, Val/Val PDMs = 20, Val/Met PDMs = 31, Met/Met PDMs = 29. PDMs differed sub-significantly from CONs in BDNF Val66Met genotype distributions ($p = 0.071$) and allele frequency ($p = 0.066$); there was a trend of excessive Met allele in PDMs than in CONs (Table 1).

Table 2 shows the odds ratios of PDM. Treat carriers (Val/Val and Val/Met) as the reference group, the odds ratio for Met/Met was statistically significant [(Met/Met PDMs x Val-carrier CONs)/(Val-carrier PDMs x Met/Met CONs)]. Treating Val/Val or Val/Met individually as the reference group, the odds ratios for Met/Met were also statistically significant. In contrast, when treating Met carriers (combining Met/Met and Val/Met as one group) or Val/Met as case (exposed group) with Val/Val as reference group, the odds ratios for Met carriers or Val/Met were not significant. These results implied that the odds of PDM...
FIGURE 1 | Examples of MEG signals at different representative time scale factors (τ) in one region (the left amygdala) from one representative subject in each group.

TABLE 1 | BDNF Val66Met (rs6265) genotype distributions and allele frequency.

| Genotype (n, %) | χ² | p | Allele frequency | χ² | p |
|----------------|----|---|-----------------|----|---|
| Val/Val        | 20 (35.0%) | 31 (38.8%) | 29 (36.6%) | 5.28 | 0.071 |
| Val/Met        |    |    |                 |    |    |
| Met/Met        |    |    |                 |    |    |
| PDM (n = 80)   |    |    |                 |    |    |
| CON (n = 76)   | 25 (32.9%) | 36 (47.4%) | 15 (19.7%) | 56.60% | 43.40% |

Significant differences were tested using chi-square tests (p < 0.05). PDM, primary dysmenorrhea patients; CON, healthy female controls; BDNF, brain-derived neurotrophic factor; Val/Val, Valine/Valine; Val/Met, Valine/Methionine; Met/Met, Methionine/Methionine.

Demographic, Pain Experiences, and Psychological Characteristics

For demographic data and menstrual features, no significant main effects of group and genotype, and no interactions of group and genotype were found (Table S2). Thus, post-hoc pairwise comparisons were not performed on demographic information. Among the three genotypes in PDMs, there were overall no differences in their menstrual pain experiences except the menstrual pain history (p = 0.006; Table S3), though post-hoc pairwise comparison revealed no significant difference in menstrual pain history between genotypes.

For psychological characteristics, there were consistently significant main effects of group but no main effects of genotype and no group × genotype interactions. PDMs reported significantly lower quality of life and higher personal emotional adjustment problems than those in CONs (Table S4). PDMs also scored higher in negative mood (depression and anxiety) and negative cognitive style to pain (pain catastrophizing) compared to CONs (Table 3).

Multiscale Sample Entropy

In this study, we focused on the interactions of BDNF Val66Met polymorphism and long-term menstrual pain. Significant group by genotype interactions were found in brain regions including the hippocampus, amygdala, insula, thalamus, putamen, superior temporal pole, supramarginal
TABLE 2 | Associations of different BDNF Val66Met genotypes with PDM.

| Reference | Case       | Odds ratio | 95% CI   | χ²     | p-value |
|-----------|------------|------------|----------|--------|---------|
| Val/Val   | Met/Met    | 2.42       | 1.03–5.69| 4.142  | 0.042*  |
| Val/Met   | Met/Met    | 2.25       | 1.02–4.93| 4.125  | 0.042*  |
| Val carrier| Met/Met    | 2.31       | 1.12–4.78| 5.248  | 0.022*  |
| Val/Val   | Met carrier| 1.47       | 0.73–2.95| 1.183  | 0.277   |
| Val/Val   | Val/Met    | 1.08       | 0.50–2.30| 0.036  | 0.849   |

*Significant odds ratios were tested using chi-square tests (p < 0.05). PDM, primary dysmenorrhea patients; CI, confidence interval; Val/Val, Valine/Valine; Val/Met, Valine/Methionine; Met/Met, Methionine/Methionine; Val carrier, Val/Val and Val/Met; Met carrier, Met/Met and Val/Met.

gyrus, superior temporal gyrus, and others (Table 4 and Figure 2).

Specifically, we tested planned comparisons of BDNF Val66Met genotypes and groups (Figure 3). For BDNF-associated regional MSE differences (between-genotype differences in each group) in CONs (Figure 3A and Table S5), MSE values were all larger in Val/Val CONs than in Val/Met CONs and Met/Met CONs. In PDMs (Figure 3A and Table S6), MSE values were also mostly larger in Val/Val PDMs than in Val/Met PDMs and Met/Met PDMs, except in the left amygdala. Val/Val PDMs manifested significant larger regional MSE in the left posterior cingulate gyrus (PCC) than in Val/Met PDMs and Met/Met PDMs, a phenomenon which was not found in the CON group. We also noticed that in both PDM and CON groups, regional MSE values in the left Heschl’s gyrus were all larger in Val/Val than in Val/Met and Met/Met groups.

For pain-associated regional MSE (between-group differences of the same genotype; Figure 3B and Table 5) in Val/Val individuals, MSE values were found to be lower in the right hippocampus and left amygdala in Val/Val PDMs than in Val/Val CONs. On the other hand, in Met/Met individuals, pain-associated regional MSE values were found to be larger in the left amygdala, left superior temporal pole, and right calcarine sulcus in Met/Met PDMs than in Met/Met CONs.

MSE profiles of six subgroups (Val/Val, Val/Met, and Met/Met in PDMs and CONs) in the right hippocampus (Figure 4), one of the most interested regions in our studies, were depicted from time scale factors τ = 1 to 100 for visual comparisons. Overall, Val/Met and Met/Met genotype groups had lower regional MSE than Val/Val group in both PDM and CON groups. At coarse time scales, the differences between Val/Val and Met carriers (Val/Met and Met/Met) were large in CONs, but such differences were diminished in PDMs. Same observations held for those in the left amygdala.

Correlations Between BDNF-Associated or Pain-Associated Regional MSE and Psychological Characteristics

Correlation results were summarized into different resting-state networks and different categories of psychological characteristics (Figure 5 and Table S7). We found that significant correlations in Met/Met group and PDM group mainly emerged in the subcortical regions (such as amygdala, hippocampus) and sensorimotor regions (such as thalamus), whereas significant correlations in Val/Val group and CON group emerged largely in the cortical regions (such as middle temporal gyrus, superior temporal gyrus, fusiform gyrus) and some of the subcortical regions (such as hippocampus). Also, after long-term menstrual pain, the correlations found in CONs were reversed or diminished in PDMs. MSE values in the amygdala showed trends towards negative correlations with both depression and anxiety scores in Met/Met CONs but were positively correlated to those in Met/Met PDMs (Figure 5A). On the other hand, MSE values in the hippocampus were also negatively correlated to depression scores in Met/Met CONs but positively correlated to those in Met/Met PDMs (Figure 5B), and were negatively correlated to anxiety scores in Met/Met CONs but positively correlated to those in Val/Val PDMs.

In PDMs, correlations between pain experiences and regional MSE mainly emerged in the Val/Val group but absent in the Val/Met or Met/Met groups. In Val/Val PDMs, pain chronification (menstrual pain history, PDM onset, menstrual pain duration) experiences were negatively correlated with MSE values. The younger the PDM onset age, or the longer the PDM history/duration, the lower the MSE in the limbic regions including the amygdala (Figure 5C), thalamus, and posterior cingulate gyrus. In contrast, pain intensity experiences (pain score, pain rating indexes) were mostly positively correlated with MSE in the thalamus (Figure 5D) and Heschl’s gyrus.

DISCUSSION

In this study, we used multiscale sample entropy analysis, a powerful tool that quantifies non-linear dynamics in time-varying signals, to investigate whether inter-subject genetic variation interacts with long-term menstrual pain experience to affect brain complexity. First, we found that BDNF Val66Met polymorphism (Met/Met homozygosity) is a potential genetic risk factor associated with primary dysmenorrhea, which is in line with previous studies (Lee et al., 2014; Wei et al., 2016b). Second, our findings indicate that long-term menstrual pain experience alters the effects of BDNF Val66Met polymorphism on brain complexity. By comparing brain complexity in females of different genotypes with or without menstrual pain, we revealed a characteristic tendency. There was considerable genotype-specific complexity differences in CONs, where Met-carrier (Val/Met and Met/Met) CONs showed extensive lower brain complexity compared to Val/Val CONs. However, the complexity differences were remarkably diminished in PDMs, implying the assaults of chronic recurrent pain on brain complexity. Third, we observed pain-associated brain complexity alterations in the limbic regions, especially the hippocampus and amygdala, in females with same BDNF Val66Met genotype.

In our recent study (Low et al., 2017), we categorized female participants according to menstrual pain experience to have a general understanding of brain complexity alterations in PDM. In this study, we further categorized all participants according to BDNF Val66Met genotypes together with menstrual...
|                      | PDM (n = 80) | CON (n = 76) | Group main effect (F) | Genotype main effect (F) | Group x Genotype interaction (F) | Between-group (p) |
|----------------------|-------------|-------------|-----------------------|-------------------------|---------------------------------|-------------------|
|                      |             |             |                       |                         |                                 |                   |
| **ANXIETY**          |             |             |                       |                         |                                 |                   |
| BAI (0–63)           |             |             |                       |                         |                                 |                   |
| Val/Val              | 7.25 (7.3)  | 3.04 (3.0)  | \(F_{(1,150)} = 17.23^{**}\) | \(F_{(2,150)} = 0.85\) | \(F_{2,150} = 0.29\) | 0.008             |
| Val/Met              | 5.94 (5.6)  | 2.94 (2.5)  |                         |                         |                                 |                   |
| Met/Met              | 7.07 (5.6)  | 4.27 (3.7)  |                         |                         |                                 |                   |
| Genotype (p)         | 0.659       | 0.511       |                         |                         |                                 |                   |
| STAI total score (40–160) |     |             |                       |                         |                                 |                   |
| Val/Val              | 82.16 (17.3) | 66.16 (9.90) | \(F_{(1,143)} = 23.14^{**}\) | \(F_{(2,143)} = 1.37\) | \(F_{2,143} = 0.74\) | < 0.0005*         |
| Val/Met              | 81.23 (16.0) | 71.62 (12.6) |                         |                         |                                 | 0.012             |
| Met/Met              | 84.31 (15.8) | 74.60 (14.6) |                         |                         |                                 | 0.049             |
| Genotype (p)         | 0.732       | 0.123       |                         |                         |                                 |                   |
| State anxiety (20–80) |             |             |                       |                         |                                 |                   |
| Val/Val              | 37.55 (8.1)  | 30.88 (5.8)  | \(F_{(1,149)} = 13.27^{**}\) | \(F_{(2,149)} = 1.43\) | \(F_{2,149} = 0.36\) | < 0.0005*         |
| Val/Met              | 37.65 (8.5)  | 33.08 (9.1)  |                         |                         |                                 | 0.063             |
| Met/Met              | 39.14 (8.2)  | 35.33 (7.7)  |                         |                         |                                 | 0.164             |
| Genotype (p)         | 0.571       | 0.199       |                         |                         |                                 |                   |
| Trait anxiety (20–80) |             |             |                       |                         |                                 |                   |
| Val/Val              | 45.05 (8.7)  | 35.28 (5.1)  | \(F_{(1,143)} = 27.82^{**}\) | \(F_{(2,143)} = 0.57\) | \(F_{2,143} = 1.08\) | < 0.0005*         |
| Val/Met              | 43.43 (8.7)  | 37.50 (9.1)  |                         |                         |                                 | 0.003*            |
| Met/Met              | 44.50 (9.2)  | 39.27 (7.9)  |                         |                         |                                 | 0.066             |
| Genotype (p)         | 0.772       | 0.244       |                         |                         |                                 |                   |
| **DEPRESSION**       |             |             |                       |                         |                                 |                   |
| BDI (0–63)           |             |             |                       |                         |                                 |                   |
| Val/Val              | 16.80 (9.1)  | 4.88 (6.8)  | \(F_{(1,150)} = 4.65\) | \(F_{(2,150)} = 2.12\) | \(F_{2,150} = 1.40\) | 0.078             |
| Val/Met              | 16.94 (13.3) | 8.03 (9.0)  |                         |                         |                                 | 0.044             |
| Met/Met              | 19.04 (9.4)  | 5.20 (7.8)  |                         |                         |                                 | 0.950             |
| Genotype (p)         | 0.614       | 0.084       |                         |                         |                                 |                   |
| **PAIN CATASTROPHIZING** |           |             |                       |                         |                                 |                   |
| PCS total score (0–52) |             |             |                       |                         |                                 |                   |
| Val/Val              | 16.80 (9.1)  | 4.88 (6.8)  | \(F_{(1,150)} = 44.47^{**}\) | \(F_{(2,150)} = 0.65\) | \(F_{2,150} = 0.50\) | < 0.0005*         |
| Val/Met              | 16.94 (13.3) | 8.03 (9.0)  |                         |                         |                                 | 0.003*            |
| Met/Met              | 19.04 (9.4)  | 5.20 (7.8)  |                         |                         |                                 | < 0.0005*         |
| Genotype (p)         | 0.570       | 0.271       |                         |                         |                                 |                   |
| Helplessness (0–16)  |             |             |                       |                         |                                 |                   |
| Val/Val              | 7.30 (4.2)  | 2.12 (2.8)  | \(F_{(1,146)} = 44.47^{**}\) | \(F_{(2,146)} = 0.65\) | \(F_{2,146} = 0.50\) | < 0.0005*         |
| Val/Met              | 7.87 (6.0)  | 3.61 (4.6)  |                         |                         |                                 | 0.001*            |
| Met/Met              | 8.46 (4.8)  | 2.33 (3.8)  |                         |                         |                                 | < 0.0005*         |
| Genotype (p)         | 0.711       | 0.422       |                         |                         |                                 |                   |
| Magnification (0–24)  |             |             |                       |                         |                                 |                   |
| Val/Val              | 2.65 (2.1)  | 1.24 (1.9)  | \(F_{(1,146)} = 21.03^{**}\) | \(F_{(2,146)} = 0.41\) | \(F_{2,146} = 0.29\) | 0.011             |
| Val/Met              | 3.10 (3.1)  | 1.48 (1.8)  |                         |                         |                                 | 0.031             |
| Met/Met              | 3.39 (2.1)  | 1.27 (1.9)  |                         |                         |                                 | 0.001*            |
| Genotype (p)         | 0.460       | 0.602       |                         |                         |                                 |                   |
| Rumination (0–12)    |             |             |                       |                         |                                 |                   |
| Val/Val              | 6.85 (4.0)  | 1.52 (2.6)  | \(F_{(1,146)} = 56.14^{**}\) | \(F_{(2,146)} = 0.07\) | \(F_{2,146} = 1.96\) | < 0.0005*         |
| Val/Met              | 5.97 (4.7)  | 2.94 (3.3)  |                         |                         |                                 | 0.006*            |
| Met/Met              | 7.18 (3.9)  | 1.60 (2.4)  |                         |                         |                                 | < 0.0005*         |
| Genotype (p)         | 0.454       | 0.145       |                         |                         |                                 |                   |

Significance main effects of group, genotype, and group x genotype interactions were tested using two-way ANOVAs (\(p < 0.01\), two-tailed); \(^*p < 0.005\); \(^{**}p < 0.00001\). Significant between-group within-genotype planned comparisons were tested using Mann-Whitney U-tests (\(p < 0.01\), two-tailed). Significant between-genotype differences in each group ("Genotype") were tested using Kruskal-Wallis H tests (\(p < 0.01\), two-tailed). Score ranges are bracketed after each item name. Data are presented as mean (SD). PDM, primary dysmenorrhea patients; CON, healthy female controls; Val/Val, Valine/Valine; Val/Met, Valine/Methionine; Met/Met, Methionine/Methionine; STAI, Spielberger state-trait anxiety inventory; BAI, Beck anxiety inventory; BDI, Beck depression inventory; PCS, pain catastrophizing scale.
pain experience. It is noted that Met allele frequencies of BDNF Val66Met polymorphism vary markedly across global populations, ranging from 0 to 72% (Petryshen et al., 2010) with higher frequency in Asian populations (more than 40%) and lower frequency in European populations (around 20%). Hence, the population genetic distribution of BDNF Val66Met in Asians allows us to recruit adequate Met-carrier participants to delineate genuine inter-subject genetic variation, as seen in our between-genotype MSE differences. We advise that combining Met carriers (Val/Met and Met/Met) or Val carriers (Val/Met and Val/Val) as one single genotype group, due to the paucity of Met/Met or Val/Val homozygotes, could overlook subtle yet informative genotype-specific changes at the brain level.

Our results demonstrated that the alterations of MSE in PDMs were majorly clustered on large time scales ($\tau = 50$–$100$), including interactions above scales 73 (Table 4), pain-associated differences above scale 78 (Table 5), BDNF-associated differences above scale 50 (Tables S5, S6), and correlations above scale 50 (Table S7). Time scales in MSE are reported to have some correspondences with signal frequencies (Mizuno et al., 2010; Courtiol et al., 2016). Coarse-graining procedure in MSE analysis resembles applying low-pass filtering or downsampling procedure to the original time-series signal. According to Nyquist-Shannon’s sampling theorem, sample entropy value at scale factor $\tau$ could reveal the irregularity of the signal under the frequency of ($f_s/\tau$)/2 Hz (Courtiol et al., 2016), where $f_s$ is the sampling frequency of the original signal. Given a sampling rate around 1000 Hz in the current study, SE values at time scales 20/50/75/100 might reveal the irregularity of the signal under 25/10/6.7/5 Hz. Our findings of large-scale MSE alterations in PDMs implicate that the resting-state neural complexity altered by the interactions of long-term menstrual pain and BDNF Val66Met polymorphism emerged approximately below theta and alpha frequency bands. Moreover, different brain regions exhibited different patterns of alterations, such as the limbic regions ($<$theta band), the sensorimotor regions ($<$alpha band), and the default mode network regions ($<$beta band). These are in line with previous studies, which reported alterations of theta oscillations at limbic regions in PDMs (Lee et al., 2017) and spectral alterations in low frequencies (theta and alpha bands) in chronic pain patients (Pinheiro et al., 2016; Ploner et al., 2017). Thus, MSE could be an important method to explore brain complexity and neural adaptability alterations.

The findings of MSE differences between BDNF Val66Met genotypes in healthy female controls support our hypothesis of genotype-specific complexity differences. Regional MSE values in Met-carrier CONs (Val/Met and Met/Met) were extensively lower at different brain regions compared with those in Val/Val individuals (Figure 3A). In healthy conditions, BDNF Val/Val homozygotes might serve a protective role on neural complexity, whereas Met allele(s) (Val/Met or Met/Met) might lead to lower neural complexity, implying a defective role of Met allele on the overall brain complexity in healthy females. One common explanation is the “neurotrophic model” that the replacement of Val by Met variant disrupts intracellular trafficking, distribution, and activity-dependent BDNF secretion at synapses in Met.
carriers (Egan et al., 2003; Chen et al., 2004). As neural complexity might reflect neuronal transients or flexible rapid transitions between neuronal microstates (Friston, 2000; Wang et al., 2018), we reason that Met carriers displayed a general loss of brain complexity compared with Val/Val homozygotes. This result is in line with previous resting-state fMRI study (Wei et al., 2016b) that healthy females with different BDNF Val66Met genotypes engaged larger variations in functional connectivity within the descending pain modulatory system. Without the influence of long-term pain experience, healthy individuals might preserve more substantial viability and flexibility in neural dynamics.

However, in PDMs, the between-genotype complexity differences in healthy females was greatly diminished (Figure 3A). This finding suggests that three genotype groups in PDMs demonstrated a less resilient brain system after the experience of long-term menstrual pain. BDNF contributes to the sensitizing capacity of the pain pathways from peripheral nociceptors, spinal level, to brain level, and is cardinally involved in central sensitization of pain (Nijs et al., 2015). Individuals with a single Val allele might “preserve,” at least to some extent, the neural complexity compared to Met carriers. In Val/Val PDMs but not in Met-carrier PDMs, the shorter the PDM history or duration was, the higher the neural complexity was preserved (close to those in Val/Val CONs) in pain-related regions (including the amygdala, thalamus, and posterior cingulate gyrus). These results suggest that the protective role of Val/Val homozygosity with respect to pain chronification is preserved in PDMs and only substitution of both alleles (as in Met/Met PDMs) might lead to loss of complexity in relatively more regions, which might reflect maladaptive neural plasticity upon long-term pain insults.

On the other hand, between-group comparisons of the same genotype (Figure 3B) revealed that neural complexity in the limbic regions (hippocampus, amygdala) was affected by long-term menstrual pain experience in a genotype-specific manner. From our recent study (Low et al., 2017), we learned that the complexity of the hippocampus and amygdala was generally lower in PDMs than in CONs. In the current study, in the hippocampus, Val/Val CONs (the reference group) had the highest regional MSE values and there was no difference among the three genotype subgroups in PDMs. This result suggests that complexity in the hippocampus might be vulnerable to long-term menstrual pain regardless of genotype. Moreover, the hippocampal MSE values were negatively correlated to anxiety and depressive scores in Met/Met CONs but positively correlated in Met/Met PDMs and Val/Val PDMs (Figure 5B), implying...
possible maladaptive neuroplasticity after the experience of long-term menstrual pain. This speculation is in accordance with the findings in animal study, in which stressful environment leads to increased anxiety-related behaviors in BDNF Met/Met mice but not in wild-type mice (Chen et al., 2004). In other words, the effect of chronic pain experience (environmental stress) might preponderantly affect the effect of BDNF Val66Met polymorphism on brain complexity.

Interestingly, in the amygdala, Met/Met PDMs showed higher regional MSE values compared with Val/Val PDMs and Met/Met CONs, suggesting that the complexity in the amygdala might be modulated differently by long-term menstrual pain in specific genotype. From the patterns of altered neural complexity observed in the amygdala and hippocampus in PDMs, we speculate the relationships lie between the amygdala, hippocampus, hypothalamus–pituitary–adrenal axis (HPA axis) system, and BDNF. Evidence from cellular to human studies indicates that pain and stress activate the HPA axis. The amygdala and hippocampus may play important yet distinct roles in the HPA axis (Smith and Vale, 2006; Weidenfeld and Ovadia, 2017). BDNF also substantially participates in the regulation of HPA axis activity (Naert et al., 2011). Under stressful environment, hippocampal atrophy and decreased BDNF secretion were found in mice (Chen Z. Y. et al., 2006). In depressive individuals,
TABLE 5 | Pain-associated (between-group) regional MSE differences.

| Brain region         | L/R Count | p-value | t-score | Cohen's d |
|----------------------|-----------|---------|---------|-----------|
| Val/Val: PDM < CON (Count = 2) |           |         |         |           |
| Limbic regions       |           |         |         |           |
| Hippocampus*         | R 1       | 0.0016  | -3.41   | -1.024    |
| Amygdala*            | L 1       | 0.0026  | -2.98   | -0.894    |
| Met/Met: PDM > CON (Count = 4) |           |         |         |           |
| Limbic regions       |           |         |         |           |
| Amygdala             | L 2       | 84.91   | 0.0007  | 3.34      | 1.063    |
| Superior temporal pole | 1       | 94      | 0.0036  | 3.00      | 0.953    |
| Visual network       | Calcarine s | R 1   | 83      | 0.0046   | 2.83     | 0.899    |

*Significant planned comparisons tested for whole brain using permutation tests (iterations = 5000, p < 0.005, two-tailed), * represents significant brain regions that also survived under stricter correction (Bonferroni-adjusted p < 0.05, two-tailed), PDM, primary dysmenorrhea patients; CON, healthy female controls; Val/Val, Valine/Valine; Met/Met, Methionine/Methionine; L, left hemisphere; R, right hemisphere; s, sulcus.

decreased level of BDNF leads to hippocampal atrophy and prefrontal cortex atrophy (Duman and Monteggia, 2006). Neuronal morphology studies in rats reveal that exposure to chronic stress leads to hippocampus atrophy but amygdala hypertrophy (Vyas et al., 2002). Both acute and chronic stress trigger opposite effects and different temporal profiles on BDNF levels in the amygdala (BLA) vs. hippocampus (CA3) in rats (Lakshminarasimhan and Chattarji, 2012). Thus, chronic stress leads to contrasting patterns of neuronal dendritic remodeling in the hippocampus and amygdala that results in dysregulation of HPA axis (Vyas et al., 2002), which are consistent with our results. Atrophy of the hippocampus causes a loss of hippocampal inhibitory control over the HPA axis, whereas hypertrophy of the amygdala causes a gain in excitatory control over the HPA axis. Our results indicate that the hippocampus might be vulnerable to long-term menstrual pain regardless of BDNF Val66Met genotypes, whereas amygdala might be affected by different BDNF Val66Met genotypes, implicating a genetic mechanism of variation in brain complexity related to chronic pain. Alternatively, we speculate that the dissociation pattern between the hippocampus and amygdala could be a manifestation of system damping; an adaptive and coping mechanism to relieve the brain from overloaded limbic information while maintaining the pain salience and harmful signal detection. These findings suggest that MSE-brain complexity could be a more sensitive measurement of neurodynamics in comparison to conventional functional connectivity observed in fMRI to reflect the central responses to brain stress (i.e., painful insults).

There is a lack of direct evidence of the association between BDNF or BDNF Val66Met polymorphism and neural complexity (as measured by MSE). A clinical case study of a single autism spectrum disorder patient reported electroconvulsive therapy-induced changes of EEG complexity and increased serum BDNF concentrations during and after therapy (Okazaki et al., 2015); yet, the authors did not offer an adequate explanation for the underlying association between BDNF level and EEG complexity. Nevertheless, we speculate a possible link between BDNF and complexity from several points based on the neurobiological functioning of BDNF (Sasi et al., 2017). First, BDNF may be a key mediator and modulator of functional synaptic plasticity, such as activation-induced long-term potentiation and long-term depression (Park and Poo, 2013; Benarroch, 2015). Second, BDNF is also reported to increase the morphology or complexity of dendritic arbors, spines, and microarchitectural integrity, thereby influencing structural plasticity (Tolwani et al., 2002; Cohen-Cory et al., 2010; Park and Poo, 2013). Third, both functional and structural plasticity reflect changes in synaptic strength that change in short and long terms. These changes are non-linear and non-stationary and are embedded in time-varying activities of neuronal populations. Therefore, quantifying the irregularity or unpredictability of these activities or synaptic dynamics might shed light on the complexity of the neural system.

From our resting-state regional MSE findings, the effects of BDNF Val or Met alleles might be far beyond simple deleterious or protective. We recognized that BDNF genetic polymorphism and BDNF protein have complex actions/regulations in the brain that could not be simplified to a single or unifying explanation based on our resting-state brain complexity study at this stage. Other confounding factors, such as age, gender (Stefani et al., 2012), environmental factors, sample size, ethnicity, and phenotype assessment, might also result in controversial findings in BDNF genetic studies (Hong et al., 2011; Notaras et al., 2015; Tsai, 2018). Moreover, it is still an ongoing debate whether the Met allele of BDNF Val66Met might serve as a deleterious role on brain structures, performances, or health (Autry and Monteggia, 2012; Tingting et al., 2014; Benarroch, 2015; Notaras et al., 2015). Part of these studies was carried out in healthy individuals to avoid possible confounding factors such as illness, medication, or genetic risk factors associated with certain diseases (Harrissberger et al., 2014). For example, a study from healthy Chinese population reported larger gray matter volume in Met/Met homozygotes (Liu et al., 2014), although the underlying mechanism remains elusive. Therefore, for individuals who experience chronic recurrent pain or long-term illnesses, the pain or stress might lead to maladaptive neural plasticity or adaptive coping strategy in different BDNF Val66Met carriers due to the activity-dependent manner of BDNF, and might come to different conclusions.

Associating genetic variation and chronic recurrent pain with brain complexity may assist in the understanding of individual neural resilience/susceptibility to pain chronication. Moreover, from the polygenic etiology view, it is implausible that BDNF is the single gene mediating menstrual pain while gene-gene interactions and epigenetic modulations have their profound contributions on chronic pain development (Denk and McMahon, 2012; Bai et al., 2017). We speculate that epigenetic modulations of long-term menstrual pain on BDNF genotypes might better explain our findings. However, to investigate the epigenetic modulations of genetic polymorphism, relevant technologies and tools are required (Weinhold, 2006), and further investigations are needed to fully comprehend the contribution of epigenetic processes to chronic pain states.

There are several limitations in this brain complexity genetic study. First, we only focused on BDNF Val66Met genetic
polymorphism; there was no information of BDNF gene expression, BDNF protein levels, or cortisol levels to test for their associations with brain complexity. Second, other pain-related genetic polymorphisms, such as BDNF rs2049046 and G-712A reported in migraine studies (Azimova et al., 2013; Sutherland et al., 2014) or OPRM1 A118G reported in PDM study (Wei et al., 2017), might also be potential candidates of genetic modulators of chronic pain-sculpted brain complexity. Finally, most neuroimaging genetic studies encounter the problem of small sample sizes compared to traditional genetic studies. Our PDM study samples were particularly limited by challenges of data acquisition of different neuroimaging modalities on the same day, rigorous inclusion/exclusion criteria, and a high exclusion rate. Nevertheless, 156 participants (80 PDMs and 76 CONs) were recruited in the present study, which was relatively large in neuroimaging studies. Future studies are invited to test the neuroimaging genetic results in brain complexity.
FIGURE 5 | Significant correlations between regional MSE and pain-related psychological characteristics. (A) Correlations between MSE in the left amygdala (τ = 91) and depression (BDI score). (B) Correlations between MSE in the right hippocampus (τ = 91) and depression score (BDI score). (C) Correlations between MSE in the left amygdala (τ = 78) and pain history. (D) Correlations between MSE in the right thalamus (τ = 58) and recalled menstrual pain index. Spearman rho (p < 0.05, two-tailed). Significant correlations are plotted as solid lines; correlations that are not significant are plotted as dashed lines. Colors: BDNF Val/Val (red), Val/Met (green), and Met/Met (black). Shapes: PDMs (circle), CONs (triangle). MSE, multiscale sample entropy; Val/Val, Valine/Valine; Val/Met, Valine/Methionine; Met/Met, Methionine/Methionine; PDM, primary dysmenorrhea patients; CON, healthy female controls; BDI, Beck depression inventory; McGill pain questionnaire; PPI, present pain index; y, year; τ, time scale factors.

CONCLUSIONS

Applying MSE analysis to time-varying MEG signals provides valuable information about neural complexity. We found that PDMs exhibited a general loss of brain complexity in pain-related regions and BDNF Val66Met polymorphism is involved in the complexity differences in a genotype-specific manner. Overall, the BDNF Val/Val homozygosity might serve as a protective role that preserves the brain complexity. Our results suggest that pain experience preponderantly affects the effect of BDNF Val66Met polymorphism on brain complexity, particularly those in the limbic circuits (hippocampus and amygdala), implicating gene-environment interaction on brain complexity.

AUTHOR CONTRIBUTIONS

L-FC and J-CH conceived and designed the experiments. IL and Y-HL performed the experiments. H-TC performed the clinical assessment. M-WL performed the genotyping. P-CK and Y-SC developed the methodology. IL, C-LT, Y-HL, and M-WL analyzed the data. J-CH provided substantial intellectual input for the work. IL, P-CK, Y-SC, L-FC, and J-CH wrote the manuscript. All authors had reviewed the paper. L-FC and J-CH approved the final submission.

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SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fnins.2018.00826/full#supplementary-material

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Conflict of Interest Statement: The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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