Diagnosed depression and sociodemographic factors as predictors of mortality in patients with dementia
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Background
Potentially modifiable risk factors for developing dementia have been identified. However, risk factors for increased mortality in patients with diagnosed dementia are not well understood. Identifying factors that influence prognosis would help clinicians plan care and address unmet needs.

Aims
To investigate diagnosed depression and sociodemographic factors as predictors of mortality in patients with dementia in UK secondary clinical care services.

Method
We conducted a cohort study of patients with a dementia diagnosis in an electronic health records database in a UK National Health Service mental health trust.

Results
In 3374 patients with 10,856 person-years of follow-up, comorbid depression was not associated with mortality (adjusted hazard ratio 0.94; 95% CI 0.71–1.24). Single patients had higher mortality than those who were married (adjusted hazard ratio 1.25; 95% CI 1.03–1.50). Patients of Asian ethnicity had lower mortality rates than White British patients (adjusted hazard ratio 0.50; 95% CI 0.34–0.73).

Conclusions
Clinically diagnosed depression does not increase mortality in patients with dementia. Patients who are single are a potential high-mortality risk group. Lower mortality rates in Asian patients with dementia that have been reported in the USA also apply in the UK.

Declaration of interest
None.

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Participants
The sample consisted of secondary mental healthcare patients from Camden and Islington National Health Service (NHS) Foundation Trust in London (UK). The Trust provides mental health services to around 440,000 individuals in two inner-city London boroughs (Camden and Islington). Patients were identified by the Clinical Record Interactive Search (CRIS) system – a platform developed to enable searches in anonymised routine electronic health records.17 The Camden and Islington CRIS database contains anonymised information on more than 117,000 mental health patients from 2008 to 2016.

Approval to use CRIS in Camden and Islington was received from the National Research Ethics Service Committee East of England – Cambridge Central (ethical approval number – 14/EE/0177). The study itself was approved by Camden and Islington NHS Trust Research Review Committee. Informed consent was not obtained as data were pseudonymised and use of the data approved by the National Research Ethics Service Committee East of England – Cambridge Central.
Measures

Mortality
In all NHS trusts, deceased patients are identified on a monthly basis via the ‘Service User Death Report’ by their NHS number, and this information is fed back into the electronic health records.

Mental health diagnoses
Diagnoses administered in UK secondary care are routinely recorded according to ICD-10 (international classification of diseases, 10th revision) criteria. We derived dementia diagnoses with ICD-10 codes F00–F03, and G30 for Alzheimer’s disease. We identified depression diagnoses in two ways. First, we derived diagnoses of major depression with the ICD-10 codes F32.0–F32.9 and F33.0–F33.9. We included depression diagnoses made 6 months before the time of dementia diagnosis (to approximate the average depressive episode duration), or anytime afterward. Second, we used natural language processing (NLP) techniques to extract indicators of depression diagnoses from free text, in those not given an ICD-10 depression code. These NLP techniques have been developed and evaluated for extracting knowledge from unstructured text data.

The main NLP technique used in CRIS has been information extraction, where unstructured text is converted into structured tables. The main NLP technique used in CRIS has been information extraction, where unstructured text is converted into structured tables. We used the NLP technique to identify any additional depression diagnoses that might have been missed by the ICD-10 codes. Our depression diagnosis variable (yes/no) consisted of ICD-10 depression diagnoses plus the additional depression diagnoses identified by NLP, among those not assigned an ICD-10 depression code. We also used a NLP app to extract data on antidepressant use 6 months before the dementia diagnosis or anytime afterward. We investigated antidepressant use (yes/no) as a separate variable because antidepressants may be prescribed to patients with dementia for indications other than depression.

Social deprivation
The Index of Multiple Deprivation (IMD) combines national census information from 38 indicators into seven domains of deprivation (income; employment; health and disability; education, skills, and training; barriers to housing and services; living environment and crime). This results in one deprivation score for 32 482 lower super output areas in England, geographical units used for the reporting of neighbourhood-level statistics. Each area has an average population of around 1500 people (about 400 households). Patients’ addresses are recorded in routinely collected clinical data. We obtained IMD scores by linking the lower super output area code of each patient’s permanent address to 2011 national data. Higher IMD scores indicate more deprived areas. IMD scores were classified into tertiles, with the lowest (least deprived) category as the reference.

Cognitive impairment
The Mini-Mental State Examination (MMSE) is frequently used by health professionals as a brief method of assessing cognitive impairment. Patients respond to 11 items, with possible scores ranging from 0 to 30. Lower scores indicate more severe cognitive impairment. MMSE scores were extracted from the CRIS database by NLP. We used the MMSE administered closest to the date of dementia diagnosis classified into tertiles, with the least severe tertile (highest scores) as the reference.

Other demographics
We extracted gender, marital status and ethnicity. We coded marital status as married, divorced (including separated), widowed or single (never married). Ethnicity was coded as White British, White other, Asian, Black or 'mixed and other' including Chinese.

Statistical analyses
All analyses were conducted with STATA, version 11. We examined longitudinal associations with univariable and multivariable Cox survival models, to calculate hazard ratios and 95% confidence intervals. Time was defined as the number of days from date of dementia diagnosis to date of death, or the end of the follow-up period. This approach accounts for differences in follow-up times between patients. Our study was restricted to clinical data collected between 2008 and 2016 because electronic records began in 2008. However, some individuals have longer follow-up periods because they were diagnosed before 2008 but their data was migrated to the electronic health record.

Missing data
There were missing data on ethnicity (8%), marital status (11%) and IMD (7%). Cumulatively, 20% of the sample were missing data on at least one of these variables. Missing data were related to our outcome and might have introduced bias, so we conducted a sensitivity analysis with multiply imputed data. We conducted multiple imputation by chained equations, using the ice command in STATA, to impute missing ethnicity, marital status and IMD data. We used all data in the analysis model to predict missing data. Data were then analysed across 25 imputed data-sets, combined by Rubin’s rules.

Results
We identified 4684 patients diagnosed with dementia. Patients without at least one MMSE were excluded from the outset, so that models could be adjusted for cognitive impairment. Fig. 1 shows the flow of these participants through the study. After including patients with data on all other study variables, the sample size was 3374.

The characteristics of the sample with complete data are presented in Table 1. Depression was identified in 247 (7%) patients. Of these, 186 (75%) were identified by ICD-10 codes and an additional 61 (25%) by the NLP app. Of those with depression, 76% had been prescribed antidepressants in the 6 months before or anytime after their dementia diagnosis. Only 18% of patients prescribed antidepressants also had a depression diagnosis. There was evidence that patients with depression had higher MMSE scores (indicating better cognitive function) at the time of dementia diagnosis. The mean MMSE score was 21.95 (s.d. 5.86) for patients with depression and 20.41 (s.d. 6.20) for patients without depression (mean difference –1.54; 95% CI –2.43 to −0.64). Patients with depression also had a lower age at diagnosis of dementia. The mean age at diagnosis of dementia was 78.37 (s.d. 8.80) for patients with depression and 81.53 (s.d. 7.84) for patients without depression (mean difference 3.16; 95% CI 2.13–4.18).

Patients were followed for a total of 10 856 person-years, with the length of follow-up ranging from 0.4 to 14.28 years (mean 3.22; s.d. 2.16). There was no evidence that length of follow-up differed according to depression. The mean follow-up time was 3.23 years (s.d. 2.23) for patients with depression and 3.22 years (s.d. 2.15) for patients without depression. By the end of follow-up, 929 patients (28%) had died. Individuals in the complete case sample had a lower mortality rate than those with missing data (hazard ratio 0.69; 95% CI 0.58–0.80).

Associations between risk factors and mortality are presented in Table 1. We found no evidence that depression was associated with...
116,936 patients in CRIS assessed for a diagnosis of dementia between 2008 and 2016

Excluded (N = 112,252)
  • No diagnosis of dementia

Diagnosis of dementia (N = 4684) assessed for MMSE

Excluded (N = 466)
  • No MMSE

Sample (N = 4218) of patients with dementia and at least one MMSE

Excluded (N = 844)
  • Missing data on one or more study variables

Final study sample (N = 3374) of patients with dementia and at least one MMSE, who had complete data on all study variables

Table 1  Characteristics of study sample and predictors of mortality (N = 3374)

| Characteristic          | N (%)       | Hazard ratio (95% CI; P value\(^a\)) | Hazard ratio (95% CI; P value\(^b\)) |
|-------------------------|-------------|-------------------------------------|-------------------------------------|
| Gender                  |             |                                     |                                     |
| Male                    | 1329 (39)   | Ref                                 | Ref                                 |
| Female                  | 2045 (61)   | 0.72 (0.63–0.82); P < 0.0001        | 0.62 (0.54–0.72) P < 0.0001         |
| Ethnicity               |             |                                     |                                     |
| White British           | 1841 (55)   | Ref                                 | Ref                                 |
| White other             | 852 (25)    | 0.84 (0.72–0.98); P = 0.026         | 0.80 (0.69–0.94); P = 0.007         |
| Asian                   | 183 (5)     | 0.47 (0.32–0.69); P < 0.0001        | 0.50 (0.34–0.73); P < 0.0001        |
| Black                   | 344 (10)    | 0.97 (0.78–1.21); P = 0.803         | 0.91 (0.73–1.14); P = 0.424         |
| Mixed and other         | 154 (5)     | 0.59 (0.40–0.97); P = 0.008         | 0.63 (0.43–0.94); P = 0.022         |
| Marital status          |             |                                     |                                     |
| Married                 | 1101 (33)   | Ref                                 | Ref                                 |
| Divorced                | 407 (12)    | 0.95 (0.76–1.19); P = 0.638         | 1.21 (0.96–1.52); P = 0.107         |
| Widowed                 | 1239 (37)   | 1.01 (0.87–1.18); P = 0.856         | 1.02 (0.86–1.21); P = 0.825         |
| Single                  | 627 (19)    | 1.12 (0.93–1.35); P = 0.233         | 1.25 (1.03–1.50); P = 0.022         |
| Age at diagnosis, years |             |                                     |                                     |
| <65                     | 98 (3)      | Ref                                 | Ref                                 |
| 66–75                   | 558 (17)    | 1.78 (0.96–3.32); P= 0.047          | 2.06 (1.10–3.86); P= 0.024          |
| 76–85                   | 1549 (46)   | 2.65 (1.45–4.82); P= 0.001          | 3.07 (1.68–5.63); P < 0.0001        |
| >85                     | 1169 (35)   | 4.50 (2.47–8.19); P < 0.0001        | 5.47 (3.00–10.03); P < 0.0001       |
| IMD tertiles            |             |                                     |                                     |
| Least deprived (1.9–26.6)| 1156 (34)  | Ref                                 | Ref                                 |
| Middle (26.7–36.4)      | 1107 (33)   | 1.07 (0.91–1.25); P= 0.401          | 1.13 (0.96–1.33); P= 0.139          |
| Most deprived (36.5–87.8)| 1111 (33) | 1.17 (1.00–1.36); P= 0.054          | 1.23 (1.05–1.44); P= 0.012          |
| MMSE tertiles           |             |                                     |                                     |
| Least impaired (24–30)  | 958 (28)    | Ref                                 | Ref                                 |
| Moderate (18–23)        | 1157 (34)   | 1.41 (1.19–1.66); P < 0.0001        | 1.43 (1.21–1.69); P < 0.0001        |
| Most impaired (0–17)    | 1259 (37)   | 1.75 (1.49–2.05); P < 0.0001        | 1.92 (1.62–2.26); P < 0.0001        |
| Depression              |             |                                     |                                     |
| No                      | 3127 (93)   | Ref                                 | Ref                                 |
| Yes                     | 247 (7)     | 0.82 (0.63–1.08); P= 0.154          | 0.94 (0.71–1.24); P= 0.645          |
| Antidepressant use      |             |                                     |                                     |
| No                      | 2321 (69)   | Ref                                 | Ref                                 |
| Yes                     | 1053 (31)   | 1.01 (0.88–1.17); P= 0.838          | 1.15 (0.99–1.33); P= 0.059          |

IMD, Index of Multiple Deprivation; MMSE, Mini-Mental State Examination; Ref, reference value.
\(a\). Univariable models.
\(b\). Multivariable model mutually adjusted for all other variables in the table.
mortality (adjusted hazard ratio 0.94; 95% CI 0.71–1.24). We found evidence that patients who reported being single had a higher mortality rate than those who were married (adjusted hazard ratio 1.25; 95% CI 1.03–1.50). Mortality rates for patients who were widowed or divorced did not differ from the mortality rate for those who were married. Mortality rates were also higher in patients living in the most deprived compared to the least deprived areas (adjusted hazard ratio 1.23; 95% CI 1.05–1.44), and in those with the most compared with the least cognitive impairment (adjusted hazard ratio 1.92; 95% CI 1.62–2.26). Mortality was also higher in patients who were older at the time of diagnosis (adjusted hazard ratio for those aged over 85 years compared with those aged under 65 years, 5.47; 95% CI 3.00–10.03). Mortality rates were lower in women than men (adjusted hazard ratio 0.62; 95% CI 0.54–0.72) and Asian patients (adjusted hazard ratio 0.50; 95% CI 0.34–0.73) compared with White British patients. There was some weak evidence of a small association between antidepressant use and mortality (adjusted hazard ratio 1.15; 95% CI 1.00–1.33). Results based on the multiply imputed sample were very similar (see Supplementary Table 1, available at https://doi.org/10.1192/bjp.2018.86).

Discussion

We found no evidence that clinically diagnosed depression was associated with mortality in people with dementia in a large mental health trust in London (UK). We found evidence that Asian patients with dementia had reduced mortality rates compared with White British patients, and that single patients with dementia had higher mortality rates compared with those who were married. Our other findings confirm associations already established in this population, that male gender, older age, more severe cognitive impairment and living in more deprived areas are associated with increased mortality rates.2,23

Strengths and limitations

To our knowledge, this is the largest study to test the association between clinically diagnosed depression and mortality in people with dementia in UK clinical care. Our clinical cohort enabled access to a large number of patients with clinically diagnosed dementia. In population-based cohorts, there is evidence that people with dementia or preclinical stages of dementia have higher drop-out rates.2,23 An advantage of routinely collected clinical cohorts is that patient attrition is reduced because of continued need for clinical care. The study setting also means that our findings are directly relevant to the management of dementia in real-world clinical practice in the UK.

Our study has several limitations. First, data obtained from routine electronic health records are not collected for research purposes, which can introduce certain issues. Patients with dementia are not routinely screened for depression in UK secondary care so it is possible that depression prevalence was underestimated in our sample. We had a relatively small number of participants with depression (n = 247). However, our total sample size was large and because we used a cohort with dementia, the mortality rate was high (>900). This large number of outcome events would have increased our statistical power. Although we found no evidence for an association between depression and mortality, it is important to note that the confidence intervals were wide. Replication in a different cohort would strengthen our conclusion that depression does not influence mortality in patients with dementia. It is also possible that depression diagnoses were assigned to patients with more severe depression, with less severe depression omitted. However, we cannot be certain that depression diagnoses were missed. Comorbidity between dementia and depression has been well documented, and clinicians are likely to be aware of this. Our investigation is also directly relevant to how depression is diagnosed in UK secondary care for dementia.

Our measurement of social deprivation relied on data obtained from the IMD. This relates to area-level census information, so may not completely capture individual socioeconomic status. We also lacked reliable information on physical health around the time of dementia diagnosis. Clinician ratings of physical health were available, but these were collected at varying time-points over the study period, and would have increased the amount of missing data.

Finally, it is possible that our results are not generalisable to other clinical settings in the UK or abroad. However, similar associations for depression and ethnicity have been reported in large studies in the USA (see below).

Depression and mortality in people with dementia

We found no evidence for an association between diagnosed depression and mortality in patients with dementia. This is consistent with several other studies of patients with dementia.2–12 However, there is good evidence for an association between depression and mortality in the general population.1–3 Many people with depression in our sample are likely to have had depressive episodes before. Given depression is associated with premature mortality, many people with depression may have died before the study or before the period of peak onset for dementia. This would mean that people with depression who survive to the main dementia risk period are no more likely to die than people without depression who also survive to this period. It is also possible that people with depression might be diagnosed with dementia at an earlier pathological stage, which could contribute to their better than expected survival. MMSE scores at the time of dementia diagnosis were higher (indicating better cognitive function) in patients with depression, and their age at diagnosis of dementia was lower.

We found some evidence of a small association between antidepressant use and mortality, consistent with another UK study.13 The prevalence of antidepressant use was 31% whereas the prevalence of depression was 7%. It is therefore possible that antidepressants are prescribed to patients with dementia for reasons other than depressed mood. This has been observed in the general population, with evidence that only 55% of antidepressant prescriptions are for depression and that they are commonly prescribed for anxiety, insomnia, pain and panic.24 It is possible that patients with dementia who were receiving antidepressants had multiple other comorbidities that could have resulted in increased mortality. Further research is therefore needed to confirm any potential association between antidepressants and mortality in patients with dementia.

Ethnicity and mortality in people with dementia

Our finding that Asian people with dementia had lower mortality rates than White British people is consistent with two large studies of patients with dementia in the USA, and one in the UK.9,13,16 Inconsistent with the studies in the USA, we found no evidence of reduced mortality rates among Black people with dementia. Higher mortality rates among White people relative to ethnic minority groups are not unique to dementia.16 The ‘mortality-cross-over’ has been well documented in the USA general population; people from ethnic minorities are more likely to die than White people until around 70 to 80 years of age. After this, mortality curves ‘cross-over’, and White people are more likely to die.16,26 One potential explanation is that, because ethnic minorities experience higher mortality rates at younger ages, a more socially and physically advantaged population remains to old age.27 To our knowledge, there are less UK data on this, but the mortality cross-over could be one explanation for our observed association.
Marital status and mortality in people with dementia

As far as we are aware, our study is the first to report that being single is associated with a higher mortality rate in patients with dementia. In contrast, we found no evidence that patients who were widowed or divorced had higher mortality rates. Our finding that being single was associated with higher mortality rates in patients with dementia should be interpreted with caution; our statistical evidence was not strong, and replication in another sample would give greater confidence in the finding. However, our evidence of increased mortality rates in single people is consistent with studies of other chronic diseases such as cancer.27 There is also evidence that, in the general population, married people tend to live longer.28 One likely explanation is that the health benefits of increased social support. For example, spouses are likely to encourage each other to lead healthy lifestyles and seek healthcare. It is also possible that more frequent social interactions increase cognitive reserve, which could slow the progression of dementia.29

Previous studies have reported that widowed and divorced people also have higher mortality rates than married people, but the effect seems consistently strongest for those who are never married.30 Our finding that being widowed or divorced was not associated with higher mortality rates could be because widowed or divorced patients had developed more or better social support networks and healthier lifestyles during the years they were married.

A potential alternative explanation for our association between single marital status and increased mortality rates is the selection of healthier individuals into marriage. However, these would need to be lifestyle and health differences that existed before marriage and were associated with risk of dementia.

Implications

Our evidence is that comorbid depression does not increase mortality rates in people with dementia. Our finding of lower mortality rates in Asian patients with dementia, if partly because of mortality cross-over, draws attention to the disparities in health outcomes that affect ethnic minority groups in the UK earlier in life. One way to address this mortality differential in patients with dementia would be to address social inequalities in health outcomes for ethnic minority groups earlier in life. Our findings also suggest that patients with dementia who are single are a higher-risk group who are more vulnerable to earlier mortality and might benefit from extra social support.

Supplementary material

Supplementary material is available online at https://doi.org/10.1192/bjp.2018.86

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Arthur Schopenhauer and psychiatry after 200 years of publication of The World as Will and Representation

Trino Baptista

This year we celebrate the 200th anniversary of publication of Arthur Schopenhauer’s text: The World as Will and Representation, which may be as pertinent for contemporary psychiatrists as it has been for artists, scientists and writers.

Schopenhauer influenced by Plato, Kant and Hinduism, conceived the will as the whole universe and a blind imperative to exist. He conceived several levels of the will: (a) an unknowable primeval level; (b) the basic forces of nature (e.g. electricity, magnetism); (c) the Platonic ideas; (d) all the inorganic and organic phenomena; and (e) deliberate human actions as the pinnacle.

By ‘representation’ Schopenhauer meant the knowledge we humans, have of the will. As all living beings, we are naturally self-centred and selfish. Besides, the Veil of Maya (a Hindu myth) standing before us, does not allow the realisation of the unity of all phenomena. Consequently, Schopenhauer had a pessimistic view of collective being and did not develop a social project. But existential (individual) salvation is possible, as a sort of illumination or ‘state of grace’, by dissociating knowledge from desire thoroughly the disinterested contemplation of art and nature, and by realising the ubiquity of suffering and of endlessly desiring. These experiences may lead to a profound inner peace and to the appearance of compassion and asceticism, these last happening as an exclusively human capacity that he stated as the ‘denial of the will’.

Schopenhauer anticipated in metaphysical terms the following issues of interest for current psychiatrists. (a) the notion of the blind and amoral will, as a precursor of some of the conceptualised features of the psychodynamic unconscious. Accordingly, he pioneered evolutionary psychology’s contemporary analysis about the relationship between the genders, particularly the prominent impact of sexuality and reproduction on cognition. (b) The pervasive human tendency for selfishness and insatiability and the value of compassion and asceticism. Redirecting these traits toward fair cooperation and healthy austerity may be considered as an ultimate goal of any comprehensive and educational psychotherapy. (c) Schopenhauer considered that we have a direct representation of the otherwise inapprehensible will, when one witnesses one’s body wishing something. I have argued elsewhere that this notion is a philosophical anticipation of the key concept of the embodied mind as discussed in contemporary cognitive sciences.

Schopenhauer’s thought is formally absent in current psychiatric education, and it is even vilified. There are many reasons for such neglect: the philosopher’s pessimism, his misogyny and his disinterest for a social project. I guess that the mistreatment of Schopenhauer’s work mainly derives from approaching his texts from secondary sources or extracts out of context. The pertinence of the philosopher’s work to contemporary culture in general and psychiatric education in particular can only be pondered when reading the original texts and placing them in their particular historical frame.