The Asymmetric Effect of Inflation on Economic Growth in Vietnam: Evidence by Nonlinear ARDL Approach

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Abstract

Low inflation and sustainable growth have been the major macroeconomic goals being pursued by every developing country, Vietnam inclusive. The effect of inflation on economic growth has been intensely analyzed by a variety of studies, but the empirical evidence more often than not remains controversial and ambiguous. One common hypothesis of previous studies is that they have assumed that the effect of inflation on growth is symmetric. The main purpose of this study is to investigate the asymmetric effect of inflation and money supply on economic growth using the Nonlinear Autoregressive Distributed Lag approach introduced by Shin, Byungchul, and Greenwood-nimmo (2013) for Vietnam over the period 1990-2017. Empirical results provide evidence that the effects of inflation on economic growth are negative and asymmetric in the long run. The impact of money supply on growth is positive in both the short-run and long-run. Accordingly, the impact of the increase in the inflation rate is bigger than the decreasing in the long-run. This different impact is significant and high inflation will destruct economic activities. As a result, the study provides empirical evidence for the authorities to plan monetary policies and control the rate of inflation to achieve sustainable economic development in the long-run.

Keywords: Inflation, Money supply, Economic Growth, Asymmetry, Vietnam.

JEL Classification Code: E31, E44, E51, E52, O42.

1. Introduction

In economic literature, the inflation rate is equal to the growth rate of nominal money minus the rate of economic growth. A crucial objective of monetary policy is to achieve high and sustainable rates of economic growth along with low and stable rates of inflation (Wulandari, Utomo, Narmaditya, & Kamaludin, 2019). Therefore, the question relating to the effects of inflation on economic growth is an intensely interesting subject of policy maker and academic. Although this subject is still debated, empirical studies are generally accepted that inflation has a negative impact on the medium and long-run growth (Bruno & Easterly, 1998; Fischer, 1993). This conclusion implies that monetary authority should maintain a low level of inflation. But, Tobin (1965) finds that inflation has a positive effect on growth. There are no effects of inflation on economic growth is a conclusion of Sidrauski (1967).

Over the last few decades, the inflation targeting regime is the crucial objective of the central banks in several countries. So, an important question is what should be the inflation target and how it is a suitable framework for the inflation target regime. According to Khan and Senhadji (2000), the levels of annual inflation between 1% and 3% is a suitable rate for industrialized countries, and between 11 and 12% for developing countries. The empirical results of Khan and Senhadji show that inflation significantly reduces growth above this threshold. At the high threshold, inflation can be destructed economic activities which reduce the variations in relative prices and, thus the negative effects of inflation on growth (Ibarra & Trupkin, 2016). Although the non-linearity of the inflation-growth nexus seems to be widely accepted, there are still controversies about: First, the level of inflation that acts as the threshold (Hung, 2017; Lam, 2015); Second, the sensitivity of this non-linear relationship to the frequency of the data, the considered framework (cross-country/time series) and the methodology used, the countries under study (developed/developing), and the existence of high inflation observation (Gillman & Harris, 2010; Lopez-Villavicencio & Mignon, 2011).
The rate of inflation is mainly depended on the amount of money supply by central banks. One common hypothesis of previous studies is that they have assumed that the effect of inflation on growth is symmetric. That means, at the low of inflation rate, an increase in the rate of inflation increases economic growth whereas a decrease in the rate of inflation decreases economic growth with the same magnitude. At the high of inflation rate, a decrease in the rate of inflation increases economic growth whereas an increase in the rate of inflation decreases economic growth with the same magnitude. There are several reasons to believe that such assumptions are not really reasonable, because: (i) Capital or labor are two main factors contributing to economic growth. In the period of economic growth, the contribution ratio of these factors must differ from the economic recession; (ii) The effects of inflation on economic growth depends on how the money is pumped into the economy and the money demand elasticity. Asymmetry effects are the different impact of the stage increases the money supply and the stage reduction of the money supply to economic growth. This asymmetric effect is the main goal in which the article wants to investigate the relationship between inflation, money supply and economic growth for Vietnam.

The remainder of the paper is as follows: Section 2 presents the theoretical background and reviews the relevant literature, section 3 shows research model, data sources and methodology, section 4 presents results interpretations and section 5 concludes and suggest several policy implications.

2. Theoretical Background and Literature Review

In the 1960s with the high-growth and low-inflation, the traditional view that inflation was destructive no longer seemed so compelling. It was the golden age of the Phillips Curve, in which economic growth and inflation were positively related in the short run. Even in the long run, Tobin (1965) find a positive effect on growth from higher inflation. There are no effects of inflation on growth is a conclusion of the study by Sidrauski (1967). According to them, when inflation was high, wealth would be reallocated away from money and onto physical capital. This conclusion is supported by the study of Wallich (1969).

Table 1: Summary of existing empirical studies

| Authors          | Countries                  | Methodology                    | Conclusions                                               |
|------------------|----------------------------|--------------------------------|-----------------------------------------------------------|
| Hayat et al. (2018) | Pakistan                   | Threshold technique            | The low inflation would be help achieve sustainable real economic growth |
| Ayres et al. (2014)   | Six regions                | OLS, FE                        | The impact of inflation targeting on real GDP is minimal overall |
| Law & Singh (2014)   | 87 countries               | Dynamic panel threshold technique | The level of financial development is beneficial to growth only up to a certain threshold; beyond the threshold level further development of finance tends to affect growth adversely |
| Bhar & Mallik (2010) | United State              | Multivariate EGARCH-M model    | Inflation uncertainty has a positive and significant effect on the level of inflation and a negative and significant effect on output growth. |
| Kiat (2008)         | 30 countries in South Africa | Ordinary Least Square          | Inflation has a negative impact on growth                  |
| Bruno & Easterly (1998) | 31 countries             | Threshold technique            | Growth falls sharply during discrete high inflation crises (40%). |
| Fountas (2010)      | 21 countries               | GARCH-in-Mean                  | Inflation uncertainty is not detrimental to output growth |
| Vaona & Schiavo (2007) | 167 countries            | Nonparametric and semiparametric estimator | At the inflation level is around 12%, it is not harmful to growth, while above it their relationship becomes markedly negative and steeper. |
| Rousseau & Yilmazkuday (2009) | 84 countries | Trilateral graphical approach | Growth is generally much lower in such high-inflation. |
| Ilyas et al. (2014)  | Pakistan                   | 2SLS technique                 | Inflation negatively and significantly affects economic growth. |
| Munir et al. (2009)  | Malaysia                   | Threshold autoregressive       | 3.89 % as the threshold value of the inflation rate above which inflation significantly retards the growth rate of GDP. Below the threshold level, there is a statistically significant positive relationship between inflation rate and growth. |
The existence of such a non-linear pattern has been confirmed by other authors, such as Burdekin, Denzau, Keil, Sitthiyot, and Willett (2004); Ghosh and Phillips (1998); Gillman and Harris (2010). From a theoretical background, the effects of inflation on growth are also mixed and depend on the way money is introduced in the models. According to Dornbusch and Frenkel (1973), if money is introduced through a pecuniary transaction cost function, the real effect of money is ambiguous. When money serves as a transaction device through the shopping, the predictions of the effects of money on growth become more clear. In endogenous growth models, the relationship between inflation and growth is accounted for via the marginal product of capital, being either physical capital (AK models) or human capital (AH models), or both. In AK models, money plays as physical capital, thus the inflation rate impacts the growth rate through its effects on the rate return of capital (Gillman & Harris, 2010; Lopez-Villavicencio & Mignon, 2011). Accordingly, inflation acts as a tax on physical capital that decreases the rate of return to capital and tends to lower economic growth. In AH models, inflation plays as a tax on human capital and also impacts the output growth rate. It leads to substitution between goods and leisure, decreasing the return to human capital, thus reduces the economic growth rate.

In sum, the effects of inflation and money supply on growth were still surprisingly ambiguous. Several empirical studies in table 1 provided evidence for this situation. As mentioned in the first section, one common hypothesis of previous studies is that they have assumed that the effects of inflation and money supply are symmetric. We suggest some reasons to believe that such assumptions are not reasonable. This gap reveals the necessity of further empirical research on the relationship between inflation and economic growth in Vietnam.

3. Data and Research Methods

3.1. Model and Data Sources

In order to investigate the asymmetric effect of inflation on the economic growth of Vietnam in the period 1990-2017 and following the previous research of (Fischer, 1993; Hwang, 2007; Tuan, 2013; Law & Singh, 2014), we specify the research model, as follows:

\[ \text{LnGDP}_t = \lambda_0 + \lambda_1 \text{INF}_t + \lambda_2 \text{MS}_t + \mu_t \quad \text{(Eq. 1)} \]

\[ \Delta \text{LnGDP}_t = \beta_0 + \beta_1 \Delta \text{LnGDP}_{t-1} + \beta_2 \Delta \text{INF} \_ \text{POS}_{t-1} + \beta_3 \Delta \text{INF} \_ \text{NEG}_{t-1} + \beta_4 \Delta \text{MS}_{t-1} + \mu_t \quad \text{(Eq. 3)} \]

The data is collected from 1990 to 2017 published by the World Bank. LnGDP is the logarithm of GDP per capita (at constant 2010 prices, unit: US dollar), INF is the rate of inflation (unit: percentage), and MS is the amount of broad money M2 per GDP (unit: percentage).

3.2. Methodology

Clearly, an estimate of Eq.1 only yields the long-run effects of exogenous variables. As mentioned in the literature review, the assumption of the symmetric impact of inflation maybe not exactly, because the increase or decrease in inflation may impact economic growth with varying degrees and orientation. In order to test the symmetry versus asymmetry effects of inflation, we follow Shin et al. (2013) and separate positive and negative changes in inflation as details:

\[ \text{INF}_t = \text{INF}_0 + \text{INF} \_ \text{POS} + \text{INF} \_ \text{NEG} \]

where: INF0 is constant (it is natural inflation). Therefore, INF_POS and INF_NEG are the partial sums of positive and negative changes in inflation. These are calculated as follows:

\[ \text{INF} \_ \text{POS} = \sum_{i=1}^{t} \Delta \text{INF}_i^+ = \sum_{i=1}^{t} \max(\Delta \text{INF}_i, 0) \]

\[ \text{INF} \_ \text{NEG} = \sum_{i=1}^{t} \Delta \text{INF}_i^- = \sum_{i=1}^{t} \min(\Delta \text{INF}_i, 0) \]

Note: \( \Delta \text{INF}_t = \text{INF}_t - \text{INF}_{t-1} \)

The long-run relationship in Eq.1 can be rewritten as follows:

\[ \text{LnGDP}_t = \lambda_0 + \lambda_1^+ \text{INF} \_ \text{POS} + \lambda_1^- \text{INF} \_ \text{NEG} + \lambda_2 \text{MS}_t + \mu_t \quad \text{(Eq. 2)} \]

Note: \( \lambda_1^+, \lambda_1^-, \lambda_2 \) are regression coefficients that present long-run impacts

In order to infer the short-run effects, we follow Pesaran, Shin, and Smith (2001); Shin et al. (2013) Bounds testing approach and rewrite Eq.2 as an Error Correction Model, as follows:

\[ \Delta \text{LnGDP}_t = \beta_0 + \beta_1 \Delta \text{LnGDP}_{t-1} + \beta_2 \Delta \text{INF} \_ \text{POS}_{t-1} + \beta_3 \Delta \text{INF} \_ \text{NEG}_{t-1} + \beta_4 \Delta \text{MS}_{t-1} + \mu_t \quad \text{(Eq. 3)} \]
where: $\Delta$ shows the first differences of the variables, $\beta_i$ are coefficients of long-run impacts, $\alpha_i$ are coefficients of short-run impacts, $\mu$ is the error.

$m_1$, $m_2$, $m_3$, $m_4$ are lag length corresponding to each variable, calculated by ARDL model following AIC, SC, HQ criteria and adjusted R-squared. And,

$$\lambda_i^* = -(\beta_i^+ / \beta_i^-), \quad \lambda_i^- = -(\beta_i^- / \beta_i^-)$$

In order to test the asymmetry effects of inflation on growth, the estimate of the nonlinear model by Ordinary Least Square technique is then used to judge four types of asymmetry. First, short-run asymmetry is established if $\lambda_i^+ \neq \lambda_i^-$ for each individual $i$. Second, the short-run impact asymmetry is established

$$\text{if } \sum_{i=0}^{m_1} \lambda_i^+ \neq \sum_{i=0}^{m_1} \lambda_i^-. $$

Third, long-run asymmetry is established if $\beta_{21}^+ \neq \beta_{21}^-$ or $\lambda_{21}^+ \neq \lambda_{21}^-$. Finally, adjustment asymmetry is captured by the pattern of dynamic multipliers. The asymmetries are tested using the Wald test. In addition to the research results are reliable, we will test the CUSUM (Cumulative Sum of Recursive Residuals) and CUSUMSQ (Cumulative Sum of Square Recursive Residuals) to check the stability of the long-run and short-run impacts.

4. Empirical Results and Discussion

4.1. Descriptive Statistics

Since 1986, Vietnam has experienced high rates of economic growth, changing the nation from a backward and underdeveloped country in the 1980s to a developing economy increasingly converging to fast-growing regional economics. In 2017, the GDP per capita of Vietnam was 1,834.65 USD (at constant 2010 prices). The monetary authority has implemented many policies to curb hyperinflation in the decade 1980s. The rate of inflation tends to decline and stabilize below 10% in the last recent years. The descriptive statistics of variables in the models are shown in Table 2.

Table 2: Descriptive statistics

| Variables | Mean | Max  | Min  | Std.Error |
|-----------|------|------|------|-----------|
| LnGDP     | 6.809| 7.515| 6.051| 0.450     |
| INF       | 11.67| 67.6 | -1.71| 16.75     |
| MS        | 70.67| 155.28| 18.91| 44.67     |

4.2. Empirical Results

4.2.1. Stationarity Test

It is well documented in the existing time series econometric literature that regression results may be spurious if the estimated variables are non-stationary and/or not cointegrated. In light of this, testing for a unit root of each series is necessary. To investigate the order of cointegration, we employed the Dickey and Fuller (1981) (ADF) test, and the Perron and Phillips (1988) (PP) test. Results of the ADF test and PP test are shown in Table 3.

Table 3: Stationarity test

| Variables | ADF test | PP test   |
|-----------|----------|-----------|
| LnGDP     | -1.271   | -1.522    |
| ΔLnGDP    | -2.716***| -3.085**  |
| INF       | -4.271***| -9.001*** |
| ΔINF      | -7.654***| -4.836*** |
| MS        | 1.278    | 1.487     |
| ΔMS       | -4.975***| -4.974*** |

***, ** and * respectively denote significance levels of 1%; 5% and 10%.

Results in Table 3 shown the INF variable is stationary at I(0). LnGDP and MS variables are stationary at I(1). No variable is stationary at I(2). Hence, the conditions to apply the NARDL of Shin et al. (2013) are satisfied.

4.2.2. Determination of Optimal Lag Length

In the NARDL model, the determination of optimal lag length is important. Thereby, the number of previous periods in which economic growth influences the current one can be identified. With the initial lag length inputted of 4, the NARDL model automatically calculates optimal lag lengths. According to AIC, SC and HC criteria, the maximum lag length is 2, in the case of Vietnam. Then, in Eq. 3, the value of $m_1$, $m_2$, $m_3$, $m_4$ is lower or equal 2.

4.2.3. Bounds Test

According to Engle and Granger (1987) on series data, there may be long-run cointegrations between variables. The long-run cointegration can be determined by F-statistics suggested by Pesaran et al. (2001).

Table 4: Results of the cointegration test

| F-Bound test for Eq.3 | Null Hypothesis: No levels relationship |
|------------------------|----------------------------------------|
| Test Statistic | Value | Signif | I(0) | I(1) |
|------------------|-------|--------|------|------|
| F-statistic      | 4.415 | 10%    | 2.37 | 3.2  |
| k                 | 3     | 5%     | 2.79 | 3.67 |
| k                 | 2.5%  | 3.15   | 4.08 |
| k                 | 1%    | 3.65   | 4.66 |
As the linear ARDL approach, Shin et al. (2013) proposed the Bounds test. Table 4 shows that F-statistic = 4.415 > Upper bound I(1) = 3.67 at the 5% significance level (this implied the null hypothesis is rejected). Thus, the results of the Bounds test reveal that there exist long-term cointegrations between LnGDP, INF, MS variables. Then Eq.3 should be estimated with the Error Correction Model.

4.2.4. Error Correction Model

There are long-run cointegrations between variables in our model, thus Eq.3 is estimated with the error correction model to determine impact coefficients in the short term. Results of the short-term impacts of inflation and money supply on economic growth are shown in Table 5.

Estimated results show that ECM(-1) = -0.1633 at the 1% significance level. This implies that economic growth is able to adjust to long-run equilibrium after each short-run “shock” created by inflation and money supply. But, the results in Table 5 also provide evidence that the asymmetric effects of inflation on growth are not clear, in the short run.

Table 5: Results of Error Correction Model

| Variable      | Coefficient | Std. Error | t-Statistic | Prob  |
|---------------|-------------|------------|-------------|-------|
| ECM(-1)       | -0.1633     | 0.0436     | -3.74       | 0.003 |
| ΔLnGDP(-1)    | 0.4837      | 0.2018     | 2.40        | 0.035 |
| ΔINF_POS      | -0.0011     | 0.0007     | -1.55       | 0.150 |
| ΔINF_POS(-1)  | -0.0013     | 0.0006     | -2.10       | 0.059 |
| ΔINF_NEG      | 0.0003      | 0.0006     | 0.55        | 0.606 |
| ΔINF_NEG(-1)  | 0.0006      | 0.0002     | 3.39        | 0.006 |
| ΔMS           | 0.0004      | 0.0002     | 2.70        | 0.013 |
| Intercept     | 1.0602      | 0.2593     | 4.09        | 0.002 |

4.2.5. Diagnostic and Stability Test

We will conduct additional diagnostic tests including the heteroskedasticity test, autocorrelation test, distribution test, stability test through examining the cumulative sum of recursive residuals (CUSUM) and cumulative sum of square recursive residuals (CUSUMSQ).

Table 6: Diagnostic test

| Type of test          | Obs * R_square                      |
|-----------------------|-------------------------------------|
| Heteroskedasticity test | 2.888 (Prob = 0.0892)              |
| Serial correlation LM test | 14.08 (Prob = 0.2284)              |
| Ramsey RESET test      | 0.767 (Prob = 0.5437)              |
| Jarque-Bera test on normality | 2.019 (Prob = 0.3643)              |

According to results in Table 6, heteroskedasticity and autocorrelation were not found in the model, and residuals were normally distributed. Figure 1 shows that both CUSUM and CUSUMSQ lines (solid lines) of Eq.3 are within the critical bounds at a significant level of 5% (dashed lines). Thus it can be concluded that Eq.3 is stable, and the estimated results are reliable for further analysis and prediction.
4.2.6. Estimated Results of Long-run Impacts

To identify the asymmetry effects of inflation on the growth of Vietnam in the 1990-2017 period, coefficients are estimated in the long run. The existence of the long-run symmetry is tested by the Wald test.

With the null hypothesis: \( H_{0LR} : \lambda^+_1 = \lambda^-_1 \)

against the alternative hypothesis: \( H_{1LR} : \lambda^+_1 \neq \lambda^-_1 \).

For checking short-run symmetry,

null hypothesis: \( H_{0SR} : \sum_{i=0}^{m^2} \alpha^+_i = \sum_{i=0}^{m^3} \alpha^-_i \)

against the alternative hypothesis:

\( H_{1SR} : \sum_{i=0}^{m^2} \alpha^+_i \neq \sum_{i=0}^{m^3} \alpha^-_i \).

If the null hypothesis is rejected, it is evidence to conclude that there is an asymmetry effect in Eq.3. According to Banerjee, Dolado, and Mestre (1998), the cumulative dynamic multiplier effects of INF_POS and INF_NEG on economic growth can be evaluated as follows:

\[
m^+_i = \sum_{i=0}^{h} \frac{\partial \text{INF}_{ret}}{\partial \text{INF }_{POS}}; m^-_i = \sum_{i=0}^{h} \frac{\partial \text{INF}_{ret}}{\partial \text{INF }_{NEG}}
\]

Note that as \( h \to \infty \) then \( m^+_h \to \lambda^+_1, m^-_h \to \lambda^-_1 \),

where \( \lambda^+_1 \) and \( \lambda^-_1 \) are the asymmetric long-run coefficients. Estimated results in Table 7 show that \( \lambda^+_1 = -0.010 \) (INF_POS) at the significance level of 5% ; \( \lambda^-_1 = 0.002 \) (INF_NEG) and not yet significant. Results of the Wald test in Table 8 show that \( W_{LR} = 9.593 \) (Prob = 0.01) this implied \( H_{0LR} \) is rejected, and \( W_{SR} = 5.803 \) (Prob = 0.035), this implied \( H_{0SR} \) is rejected. These results provide evidence to conclude the effect of inflation on economic growth is asymmetry. Accordingly, there is a negative impact on growth from inflation in the long run.

Table 7: Estimated results of long-run impacts

| Variables | Coefficient | t_statistic | Prob  |
|-----------|-------------|-------------|-------|
| INF_POS   | -0.010      | -7.060      | 0.022 |
| INF_NEG   | 0.002       | 0.624       | 0.446 |
| MS        | 0.016       | 30.67       | 0.000 |
| Long-run asymmetry test | Short-run asymmetry test | | |
| INF       | \( W_{LR} = 9.593 \) (Prob = 0.010) | \( W_{SR} = 5.803 \) (Prob = 0.035) |

5. Conclusion and Policy Implications

Using annual data in the period 1990-2017, by applying Nonlinear Autoregressive Distributed Lag approach proposed by Shin et al. (2013) this study affirms two following main points:

(i) There is a strong statistics evidence to conclude that the impact of inflation on economic growth is asymmetry in the long run. There is a negative impact of inflation on economic growth in the case of Vietnam.

(ii) Money supply has a positive impact on growth in both the short-run and long run. Based on empirical results, the authors propose some policy implications, as follows:

Firstly: The monetary authorities need to rotate the expanded monetary policy and the shrinking monetary policy. Because increasing the money supply will help promote growth rate, but the extreme money supply policy maybe harms growth.

Secondly: Inflation has negative impacts on economic growth in the case of Vietnam, and high inflation will destruct economic activities. In order to achieve high and sustainable rates of growth, the Government should maintain low and stable rates of inflation.

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