A 23-year study of mortality and development of co-morbidities in patients with obesity undergoing bariatric surgery (laparoscopic gastric banding) in comparison with medical treatment of obesity

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Abstract

Background and aim: Several studies have shown that bariatric surgery reduces long term mortality compared to medical weight loss therapy. In a previous study we have demonstrated that gastric banding (LAGB) is associated with reduced mortality in patients with and without diabetes, and with reduced incidence of obesity co-morbidities (cardiovascular disease, diabetes, and cancer) at a 17 year follow-up. The aim of this study was to verify at a longer time interval (23 years) mortality and incidence of co-morbidities in patients undergoing LAGB or medical weight loss therapy.

Patients and methods: As reported in the previous shorter-time study, medical records of obese patients [body mass index (BMI) > 35 kg/m² undergoing LAGB (n = 385; 52 with diabetes) or medical treatment (controls, n = 681; 127 with diabetes), during the period 1995–2001 (visit 1)] were collected. Patients were matched for age, sex, BMI, and blood pressure. Identification codes of patients were entered in the Italian National Health System Lombardy database, that contains life status, causes of death, as well as exemptions, prescriptions, and hospital admissions (proxies of diseases) from visit 1 to June 2018. Survival was compared across LAGB patients and matched controls using Kaplan–Meier plots adjusted Cox regression analyses.

Results: Final observation period was 19.5 ± 1.87 years (13.4–23.5). Compared to controls, LAGB was associated with reduced mortality [HR = 0.52, 95% CI 0.33–0.80, p = 0.003], significant in patients with diabetes [HR = 0.46, 95% CI 0.22–0.94, p = 0.034], borderline significant in patients without diabetes [HR = 0.61, 95% CI 0.35–1.05, p = 0.076]. LAGB was associated with lower incidence of diabetes (15 vs 75 cases, p = 0.001), of CV diseases (61 vs 226 cases, p = 0.009), of cancer (10 vs 35, p = 0.01), and of renal diseases (0 vs 35, p = 0.001), and of hospital admissions (92 vs 377, p = 0.001).

Conclusion: The preventive effect of LAGB on mortality is maintained up to 23 years, even with a decreased efficacy compared with the shorter-time study, while the preventive effect of LAGB on co-morbidities and on hospital admissions increases with time.
Introduction

Patients with obesity undergoing bariatric surgery have a longer life expectancy than patients receiving medical treatment of obesity. Several papers [1–8], analyzed in two meta-analyses [9, 10], have shown lower long-term mortality with bariatric surgery in comparison with nonsurgical controls; further, reduced mortality is observed in patients with and without diabetes [1, 4, 11, 12]. In addition, bariatric surgery improves quality of life in morbid obesity [13], is associated with lower development of medical complications of obesity, reduced frequency of co-morbidities, improved cardiovascular (CV) risk profile [14–20], and is cost-effective in the management of obesity [21, 22]. The majority of studies has been performed through well established restrictive or mixed techniques [gastric banding (LAGB), vertical banded gastroplasty (VGB), roux-en-y gastric bypass (RYGB)], but recent studies have shown that laparoscopic sleeve gastrectomy (LSG) [23], as well as malabsorptive surgery [biliointestinal bypass (BIBP) and biliopancreatic diversion (BPD)] is associated with reduced mortality and lower development of obesity related co-morbidities, compared to medical weight loss treatment of obesity [24].

No intermediate evaluation of clinical and metabolic effects of bariatric surgery, in comparison with medical treatment of obesity, has appeared in previous studies evaluating long-term mortality, so that reduced mortality seems an all-or-none effect, with no mechanistic explanation for the reduced mortality.

In a previous retrospective study we have shown that, up to 17 years, LAGB is associated with reduced mortality in patients with and without diabetes, and with reduced incidence of diabetes and cardiovascular diseases [11]. This was the longest follow-up study, with no patient lost to follow-up; we also hypothesized that a longer follow-up was required to establish if the effects of LAGB were maintained or even made more significant through a prolonged observation, or whether the effects of LAGB vanished, also because of the process of aging.

The aim of this retrospective study was to extend the follow-up period observation of the previous study up to 23 years. In addition, we had the opportunity to compare the intermediate clinical and metabolic effects of bariatric surgery and of medical treatment of obesity, thus evaluating a possible mechanistic explanation for the reduced mortality.

Methods

Patients and study protocol

The participating institutions offer surgical and medical treatment of obesity. The institutions belong to the LAGB10 study group [11], a spontaneous network of physicians and surgeons working with bariatric surgery in the Lombardy Region (Italy); LAGB has been performed here since 1995, according to NIH guidelines [25]. The specific study protocol was approved by four Ethics Committees in 2012, after the initial protocol had been approved in 1995, in 2002, and in 2006. Being a retrospective study, informed consent was obtained from all individual participants included in the study who could be reached by interview, phone or letter. The details of the protocol have been previously published [11]. Briefly, we considered all patients with obesity (BMI > 40 kg/m² alone or BMI > 35 kg/m² in the presence of co-morbidities) aged 18–65 years, seeking medical advice and referred to the outpatients obesity clinics during the period 1995–2001, (first visit) undergoing thereafter LAGB, or medical weight loss treatment. After evaluation of indications and contra-indications, patients were offered LAGB; several patients declined the offer, mainly because of reluctancy, lack of knowledge of the possible benefits, fear of surgery and of surgical complications, inability or unwillingness to comply with the anticipated change of lifestyle habits or with the program of scheduled visits. Patients who declined surgery for any reason, but agreed to be followed-up during medical treatment, were considered controls. All surgery and nonsurgical patients were treated with diet, and received standard care (education on eating behaviors, advice on diet and exercise, plus drug treatment for diabetes and hypertension when present). At least initially, all patients were evaluated under basal conditions and at 3-month intervals with measurement of body weight and assessment of food intake through review of diet diaries; their suggested diet was between 1000 and 1200 kcal/day for women and men (22% protein, 29% lipids, and 49% carbohydrates), respectively, with the aid of a dietitian. From the medical records, birthdate and age, baseline anthropometric data (height, weight, BMI) systolic and diastolic blood pressure, heart rate, metabolic data (fasting blood glucose, glycated hemoglobin [HbA1c (%)], total cholesterol, HDL-, and LDL-cholesterol, triglycerides, aspartate transferase [AST], alanine transferase [ALT], creatinine and eGFR [modified diet in renal disease calculation.
equation] [26]), current medical treatments, clinical evidence of coronary heart disease (CHD), retinopathy, were derived and tabulated. From the medical records it was also possible to evaluate later visits and lab examination, when present. Diagnosis of diabetes (type 2 diabetes) was established as already reported [27, 28], and diagnosis of coronary heart disease (CHD) was based on medical records.

Procedures

Patients were identified through personal identification codes; codes were entered the Regional Lombardy Administrative Database, and it was possible to ascertain whether patients were alive, were dead, or had moved to other regions. The National Health System (NHS) covers more than 95% of all hospital admissions, medical and surgical procedures and medical expenses of citizens [29] (Italian Survey 2012). The Regional Lombardy Administrative Database contains since 1988 all pertinent data of all citizens, and this makes life status a clear finding, independently of participation in studies and of loss to follow-up. In particular, the Lombardy database collects several information, including (1) an archive of residents who receive NHS assistance, reporting demographic and administrative data; (2) a database on diagnosis at discharge from public or private hospitals of the region; (3) a database on outpatient drug prescriptions reimbursable by the NHS; and (4) a database on outpatient visits, including visits in specialist ambulatory care and diagnostic laboratories accredited by the NHS. For each patient, these databases are linked through a single identification code.

In the Italian National Health System development of chronic diseases (diabetes mellitus, liver and cardiovascular diseases, selected thyroid, renal, and lung diseases) yields the right to exemption from medical charges (exemptions), that means life-long free prescriptions and examinations for the above diseases. Therefore, together with hospital admissions, exemptions were considered a proxy of development of chronic diseases. For each patient, exemptions and hospital admissions after first visit were identified and dated. Through registries of surgeons and the Regional Lombardy Administrative Database it was also possible to retrieve patients who had removal of LAGB and/or new bariatric surgery procedures. Through the health districts (ASL) patients belonged to, it was possible to track causes of death, and nature of hospital admissions and of exemptions. Data from health districts were cross-checked with data from the Lombardy Database, to rule out inconsistencies and possible delays in transcriptions. This procedure has already been employed and validated in previous studies in Lombardy, Italy [11, 30]. The limit date of June 30, 2018 was established for all patients for deaths, admissions, and exemptions. Causes of death, as well as exemptions and hospital admissions were coded according to ICD-10 codes. Full details of the procedures are reported elsewhere [11, 24, 30].

Outcomes

Death rate and cause of death among patients with diabetes (surgical vs nonsurgical) and among patients without diabetes (surgical vs nonsurgical); exemptions and hospital admissions among patients with and without diabetes (surgical vs nonsurgical). Analysis of survival and of other outcomes was carried out on an intention-to-treat basis, with no consideration for LAGB removal.

Statistical analysis

Data are shown as average values (± SD) for continuous variables or absolute numbers and frequencies for discrete variables. Continuous variables were compared with the Student’s t-test. Frequencies were compared with the Fisher exact test. Surgical and nonsurgical patients were matched (with and without diabetes separately) with no attempt to match patients of the whole cohort. Group matching was made for sex, BMI (± 5 kg/m²), age (± 10 years), for systolic (± 5 mmHg), and diastolic (± 5 mmHg) blood pressure. The median age of matched patients was 42 years, and the mean ages were 31.8 ± 6.43 and 51.8 ± 5.89, respectively. The proportion of dead patients was plotted through Kaplan–Meier curves, and differences in survival among subgroups were tested by the log-rank test. A multivariable analysis of risk factors for mortality was performed (Cox proportional hazards model), and used to plot Kaplan–Meier curves for surgery versus nonsurgical patients; age, median age, presence of diabetes, sex, systolic blood pressure, eGFR, and presence of CHD were entered a priori. Proportionality among the survival rates and attributable factors in the Cox model was assessed by plotting the log [− log (survival function)] versus time. Statistical analyses were performed with STATA 12.0 for MacIntosh.

Power calculation and sample size

Being a retrospective study, power calculation and sample size were only calculated to understand if the study was meaningful. Due to previous papers dealing with long-term prevention of mortality, showing effectiveness of about 50% in comparison with non-surgery subjects [9, 10], given a power = 80% and an alpha error 0.05, it was calculated that 500 surgery subjects with 30 fatal events and 1000 nonsurgical subjects with 90 fatal events were required to detect significant differences in the outcomes [31, 32]. Similarly, given the high efficacy of bariatric surgeries in the long-term prevention of diabetes
and of cancer, [33–35], we estimated that the occurrence of 100 exemptions in 500 bariatric surgery subjects and 300 exemptions in 1500 subjects undergoing dietary and medical treatment would be required to detect significant differences in the outcomes between the two groups [31, 32]. This manuscript was prepared following the guidelines of the STROBE statement [36] (Additional file 1).

Results

The details of patients in the study were already published in a previous publication [11], and now appear in Additional file 2: Table S1. Observation period was 19.5 ± 1.87 years (13.34–23.5). Mortality rate was 2.6, 6.6, 10.1, and 13.4% in controls at 5, 10, 15, and 20 years, respectively; mortality rate was 0.8, 2.5, and 3.1, and 7.4% in LAGB patients at 5, 10, 15, and 20 years, respectively.

Figure 1 shows crude mortality curves in patients receiving LAGB as compared to controls receiving medical weight loss therapy, and Fig. 2a and b show crude mortality curves in patients without and with diabetes, respectively. The reduced mortality in surgical vs nonsurgical patients was significant in patients aged > 42 years, not significant in patients aged < 42 years. Table 1 shows causes of death in the whole cohort in the original study and in the follow-up study; causes of death were similar in the two observation periods, and the comparison between surgical vs nonsurgical patients had a reduced level of significance in the follow-up period, in agreement with the reduced overall effect on prevention of mortality.

Table 2 compares the 17 year and the 23 year effects of LAGB as opposed to medical weight loss therapy; the effect on reduced mortality decreases with time, while the effect on prevention of co-morbidities and the effect on prevention of hospital admissions increases with time.

Table 3 shows the clinical and metabolic effects of LAGB and medical weight loss therapy. The interval between baseline and follow-up data was 4.9 ± 3.63 years (mean ± SD), with no differences between surgery and nonsurgical patients. The effects were clearly different, with the noticeable exceptions of cholesterol (total, LDL-, and HDL-cholesterol).

Table 4 shows univariate and multivariate analysis of risk factors for mortality in the current study as
compared with the original study, and indicates that risk factors considered in the original study maintained their value in the follow-up study.

**Discussion**

To our knowledge, this study represents the longest follow-up evaluation of patients undergoing LAGB, a bariatric surgery, in comparison with patients receiving
weight loss medical treatment. With its up to 23 years duration of observation, this study adds about 6 years to our previous study, in the same cohort, studied in the same way. The main finding, in comparison with our previous study [11], is the somehow reduced effect on prevention of long-term mortality in comparison with our previous study; in contrast, the preventive effect of surgery on incident diseases increases, and
the preventive effect of surgery on hospital admissions increases. Therefore, it appears that the beneficial effect of LAGB continues up to 23 years, even with some differences; the effect on mortality decreases, even if it is still significant, while the effect on general health status continues, and increases. Overall, as recently confirmed

| Group              | Original study | p     | Follow-up study | p     |
|--------------------|----------------|-------|----------------|-------|
|                    | Surgery (N)    |       | Nonsurgical (N)|       |
| CVD                | 5              | 0.001 | 8              | 0.029 |
| Total non-CVD      | 7              | 0.001 | 18             | 0.019 |
| Cancer             | 7              | 0.016 | 13             | 0.033 |
| Liver              | 0              |       | 2              | NS    |
| Lung               | 0              |       | 1              | NS    |
| Infection          | 0              |       | 1              | NS    |
| Endocrine          | 0              |       | 0              | NS    |
| External           | 0              |       | 1              | NS    |
| Total              | 12             | 0.001 | 26             | 0.001 |

**Table 1** Causes of death in surgery and nonsurgical patients during the original study (observation period 13.9 ± 1.87 years, mean ± SD, 10) and in the follow-up study (observation period 19.5 ± 1.88 years)

| Group              | Original study | p     | Follow-up study | p     |
|--------------------|----------------|-------|----------------|-------|
|                    | Surgery (N)    |       | Nonsurgical (N)|       |
| Mortality          | HR = 0.35, 95% CI 0.19–0.65 | 0.001 | HR = 0.52, 95% CI 0.33–0.80 | 0.003 |
| In non-DM          | HR = 0.42, 95% CI 0.19–0.97 | 0.041 | HR = 0.61, 95% CI 0.35–1.05 | 0.076 |
| In DM              | HR = 0.34, 95% CI 0.13–0.87 | 0.025 | HR = 0.46, 95% CI 0.22–0.94 | 0.034 |
| Below median age   | HR = 0.69, 95% CI 0.18–2.68 | 0.586 | HR = 1.98, 95% CI 0.76–5.14 | 0.162 |
| Above median age   | HR = 0.29, 95% CI 0.14–0.58 | 0.001 | HR = 0.38, 95% CI 0.23–0.65 | 0.001 |

**Table 2** Comparison of mortality (HR with 95% CI), incident diseases, and hospital admissions in surgery and nonsurgical patients during the original study (observation period 13.9 ± 1.87 years, mean ± SD, 10) and in the follow-up study (observation period 19.5 ± 1.88 years)

| Incident diseases | Original study | p     | Follow-up study | p     |
|-------------------|----------------|-------|----------------|-------|
| Diabetes          | 15             | 0.018 | 15             | 0.001 |
| Arterial hypertension | 42        | 0.028 | 47             | 0.001 |
| CVD               | 10             | 0.025 | 14             | 0.009 |
| Renal diseases    | 0              |       | 0              | 0.001 |
| Liver diseases    | 8              |       | 8              | NS    |
| Cancer            | 4              |       | 10             | 0.01  |
| Lung diseases     | 4              |       | 4a             | NS    |
| Metabolic diseases| 4              |       | 4a             | NS    |
| Total             | 87             | 0.001 | 102            | 0.001 |

| Hospital admissions | Original study | p     | Follow-up study | p     |
|---------------------|----------------|-------|----------------|-------|
| Diabetes            | 14             |       | 14             | 0.001 |
| CVD                 | 34             |       | 35             | 0.001 |
| Liver disease       | 4              |       | 6              | 0.028 |
| Renal diseases      | 0              |       | 0              | NS    |
| Cancer              | 9              |       | 10             | 0.005 |
| Lung diseases       | 5              |       | 5a             | NS    |
| Metabolic diseases  | 5              |       | 5a             | NS    |
| Muscular and bone diseases | 17 | 0.04 | 17a | 0.001 |

* When no new incident diseases of hospital admissions were recorded, data from the original study are indicated
by recent 4–5 year studies performed through various surgical techniques (LGB, RYGB, LSG) [23], our data confirm that bariatric surgery is associated with lower mortality compared to medical weight loss treatment [9, 10]; also prevention of co-morbidities, especially diabetes mellitus, is possible for prolonged periods [27, 33, 37, 38].

A greater effect on mortality in patients with diabetes than in patients without diabetes has already been reported [12], leading to the interpretation that the benefit is greater in more compromised patients. There are no explanations for these differences, though it seems reasonable to assume that the aging process dilutes the preventive effect of LAGB on mortality. In the Swedish obesity study (SOS study) [37] it was observed that the preventive effect of surgery on incident co-morbidities increases with duration of follow-up (from 2 to 10 years); our data support these findings, even though the observation periods of the two studies are quite different. However, we observed that the effect of surgery depends on age, i.e. it is significant for patients above median age (42 years in this cohort), not in younger patients. This confirms what was already observed by us and by others, using different bariatric

| Table 3 Variables evaluated at baseline and follow-up (4.9 ± 3.63 years) |
|------------------------------------------------------------|
|                | Surgery (n = 154) | p            | Nonsurgical (n = 360) | p            |
|                | Baseline Follow-up |             | Baseline Follow-up |             |
| Age (years)    | 41.0±10.13        | –            | 42.2±12.94        | –            |
| BMI (kg/m²)    | 42.7±4.62         | 36.7±5.24    | 0.0001            | 39.1±5.27    | 39.1±6.16    | 0.5047 |
| Blood glucose (mg/dL) | 108.6±39.36 | 97.6±32.09 | 0.0001 | 103.1±27.40 | 103.6±30.47 | 0.4452 |
| HbA1c (%)      | 6.0±1.35          | 5.5±1.01     | 0.0001            | 7.6±2.51     | 6.7±1.64     | 0.1055 |
| Systolic BP (mmHg) | 133.8±14.59 | 127.2±12.38 | 0.0001 | 132.4±13.36 | 132.8±8.76  | 0.0900 |
| Diastolic BP (mmHg) | 83.9±9.55     | 77.8±8.68   | 0.0001            | 78.65±9.48   | 78.8±10.43  | 0.5203 |
| Heart rate (bpm) | 78.2±5.95      | 55.7±28.89  | 0.0304            | 71.2±6.46    | 75.8±10.56  | 0.1280 |
| EGFR (mL/min/1.73 m²) | 107.3±29.07 | 99.3±23.34  | 0.0010            | 85.4±21.54   | 90.2±18.83  | 0.0811 |
| Cholesterol (mg/dL) | 207.7±43.77  | 203.2±36.77 | 0.0685            | 210.5±32.68  | 198.8±37.43 | 0.0272 |
| LDL-cholesterol (mg/dL) | 130.6±40.50 | 124.4±32.60 | 0.0250     | 135.4±35.07 | 119.3±34.81 | 0.0214 |
| HDL-cholesterol (mg/dL) | 50.1±13.52 | 54.6±13.87  | 0.0001            | 48.4±11.82   | 51.29±12.29 | 0.0298 |
| Triglycerides (mg/dL) | 140.4±76.49 | 122.2±66.93 | 0.0010 | 139.2±68.34 | 130.8±52.72 | 0.1735 |
| AST (U/L)       | 23.7±11.86       | 21.5±8.55    | 0.0146            | 23.6±10.41   | 24.5±7.26   | 0.2830 |
| ALT (U/L)       | 31.3±21.14       | 24.2±15.39   | 0.0001            | 32.9±27.64   | 30.4±15.19  | 0.2568 |

Mean±SD
BMI: body mass index, HbA1c: glycated hemoglobin, EGFR: estimated glomerular filtration rate (mL/min/1.73 m²), AST: aspartate transaminase, ALT: alanine transaminase

| Table 4 Univariate and multivariable analysis of risk factors for mortality (Cox proportional hazards model) in the whole sample |
|------------------------------------------------------------|
|                | HR       | S.E.     | Z        | p         | 95% CI        |
| Univariate analysis                                     |
| Surgery        | 0.52 (0.35) | 0.12 (0.11) | −2.94 (−3.33) | 0.003 (0.001) | 0.33–0.81 (0.19–0.65) |
| Age > 42 years  | 5.53 (7.15) | 1.45 (2.43) | 6.52 (5.81) | 0.001 (0.001) | 3.31–9.26 (3.68–13.91) |
| Female sex     | 0.53 (0.39) | 0.10 (0.09) | −3.27 (−4.02) | 0.001 (0.001) | 0.36–0.78 (0.25–0.62) |
| Coronary heart disease | 4.98 (4.67) | 1.35 (1.52) | 5.94 (4.73) | 0.001 (0.001) | 2.93–8.47 (2.47–8.86) |
| Diabetes       | 3.94 (5.71) | 0.74 (1.31) | 7.29 (7.61) | 0.001 (0.001) | 2.73–5.70 (3.34–8.94) |
| Multivariable analysis                                  |
| Surgery        | 0.51 (0.41) | 0.12 (0.13) | −2.95 (−2.82) | 0.003 (0.005) | 0.33–0.80 (0.22–0.76) |
| Age > 42 years  | 4.31 (4.35) | 1.21 (1.57) | 5.21 (4.08) | 0.001 (0.001) | 2.49–7.45 (2.15–8.82) |
| Female sex     | 0.53 (0.39) | 0.10 (0.09) | −3.25 (−4.10) | 0.001 (0.001) | 0.36–0.78 (0.25–0.61) |
| Coronary heart disease | 2.83 (2.51) | 0.78 (0.83) | 3.77 (2.75) | 0.001 (0.006) | 1.65–4.87 (1.31–4.81) |
| Diabetes       | 2.65 (3.11) | 0.53 (0.75) | 4.86 (4.69) | 0.001 (0.001) | 1.79–3.93 (1.93–4.99) |

In brackets values observed in the original study [10]
techniques [5, 8, 11, 39]; in the SOS Study, patients aged < 37 years were intentionally excluded because of the low mortality of patients with obesity in young age [4].

This study has strengths and limitations; the main strength lies in the prolonged observation period of the same cohort, evaluated with the same approach; also, due to the methods employed, no patient was lost to follow-up. In addition, we had detailed description of causes of death of all patients, of incident diseases, of hospital admissions. More, we had the possibility to observe clinical and metabolic variables in a fair proportion of patients after a mean period of 5 years, and we could observe a significant different effect of surgery vs medical weight loss treatment. Obesity, and especially visceral obesity, favor development of cardiovascular disease in type 2 diabetes [40], and both type 2 diabetes and obesity predict all-cause mortality [41, 42]; the present results indicate that LAGB, able to induce weight loss and to prevent diabetes, prevents mortality through improvement of the general health status [43]. Finally, as reported above, we confirmed a significant age-related effect on prevention of mortality, in agreement with previous studies [5, 8, 11, 39].

The main limitation lies in the retrospective nature of the study; the second limitation is that the study was not randomized, but at the time this study was conceived, randomization was deemed unethical, so that prospective studies could not be performed. The fact that several patients refused surgery for multiple reasons might represent a selection bias; however, it should be emphasized that in the years 1995–2001 evidence of benefits of bariatric surgery were still limited. Also, during the first 5 years there were 4 deaths (1 above median age) in the surgery group and 18 deaths (17 above median age) in the nonsurgical group (NS); we have no explanation for a higher number of early deaths in both groups is higher than in previous papers [10], but differences in different cohorts can occur. The fourth limitation is in the sample size. The fifth limitation is represented by the fact that the use of of LAGB is declining, so that some people argue LAGB should be abandoned; actually, LAGB is still performed in a significant proportion of patients with obesity. The last limitation is that our results can not be generalized to all bariatric procedures, also because there are no studies of similar duration performed with other bariatric techniques.

Conclusion

The preventive effect of LAGB on mortality is maintained up to 23 years, even with a decreased efficacy, while the preventive effect of LAGB on incident diseases and on hospital admissions increases with time. These data indicate that the beneficial effects of LAGB is long lasting.

Additional files

Additional file 1. Strobe statement.
Additional file 2: Table S1. Subjects in the study.

Abbreviations

ALT: alanine transferase; AST: aspartate transferase; BIBP: biliointestinal bypass; BPD: bilipancreatic diversion; BMI: body mass index; CV: cardiovascular; CI: confidence intervals; CHD: coronary heart disease; RYGB: roux-en-y gastric bypass; eGFR: estimated glomerular filtration rate; LAGB: gastric band; HbA1c (%): glycated hemoglobin; HR: hazard ratio; ASL: health districts; NHS: National Health System; LSG: laparoscopic sleeve gastrectomy; SOS study: Swedish obesity study; VGB: vertical banded gastroplasty.

Authors’ contributions

AEP planned the research, contributed to discussion, wrote the manuscript; ASZ searched data, prepared the database, contributed to analysis, contributed to discussion; MF searched data, prepared database, contributed to analysis; ET performed statistical analysis, contributed to discussion; AS searched data, prepared database, contributed to discussion; EM searched data, contributed to database, contributed to discussion; CO searched data, contributed to discussion; GM searched data, contributed to discussion, edited the manuscript; FF searched data, contributed to discussion, edited the manuscript. All authors read and approved the final manuscript.

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Competing interests

The authors declare that they have no competing interests.

Availability of data and materials

Trial registration Trial registration does not apply, since this is a retrospective study. Data are available on request.

Consent for publication

All authors agree with publication.

Ethics approval and consent to participate

The specific study protocol was approved by four Ethics Committees in 2015. Being a retrospective study, informed consent was obtained from all individual participants included in the study who could be reached by interview, phone or letter.
References

1. MacDonald KG Jr, Long SD, Swanson MS, Brown BM, Morris P, Dohm GL, Pories WJ. The gastric bypass operation reduces the progression and mortality of non-insulin-dependent diabetes mellitus. J Gastrointest Surg. 1997;1:213–20.

2. Christou NV, Sampalis JS, Liberman M, Look D, Auger S, McLean AP, et al. Cardiovascular events after bariatric surgery in obese patients with type 2 diabetes. Diabetes Care. 2012;35:263–7.

3. Flum DR, Dellingler EP. Impact of gastric bypass operation on survival: a population-based analysis. J Am Coll Surg. 2004;199:543–51.

4. Sjöström L, Narbro K, Sjöström CD, Karason K, Larsson B, Wedel H, Lystig T, et al. Laparoscopic adjustable gastric banding prevents type 2 diabetes and induces their remission in morbid obesity: a 4-year case–controlled cohort study. Diabetes Obes Metab. 2017;19:1223–32.

5. Buchwald H, Estok R, Fahrbach K, Banel D, Jensen MD, Pories WJ, et al. Long-term cardiovascular risk and coronary events in morbidly obese patients treated with laparoscopic gastric banding. Surg Obes Relat Dis. 2014;10:112–20.

6. Adams TD, Gress RE, Smith SC, Halverson RC, Simper SC, Rosamond WD, Lamonte MJ, Stroup AM, Hunt SC. Long-term mortality after gastric bypass surgery. N Engl J Med. 2007;357:741–52.

7. Sowemimo OA, Yood SM, Courtney J, Moore J, Huang M, Ross R, McMillian U, Qo O, Reinhold RB. Natural history of morbid obesity without surgical intervention. Surg Obes Relat Dis. 2007;3:73–7.

8. Peeters A, D’Brien PE, Laurie C, Anderson M, Wolfe R, Flum D, Macninis RJ, English DR, Dixon J. Substantial intentional weight loss and mortality in the severely obese. Ann Surg. 2007;246:1028–33.

9. Pontiroli AE, Morabito A. Long-term prevention of mortality in morbid obesity through bariatric surgery. A systematic review and meta-analysis of trials performed with gastric banding and gastric bypass. Ann Surg. 2011;253:484–7.

10. Cardoso L, Rodrigues D, Gomes L, Carrilho F. Short- and long-term mortality of non-insulin-dependent diabetes mellitus. J Gastrointest Surg. 1997;1:213–20.

11. Pontiroli AE, Morabito A, Babel D, Jensen MD, Pories WJ, et al. Bariatric surgery: a systematic review and economic evaluation. Health Technol Assess. 2009;13:1–190.

12. Keating CL, Dixon JB, Moodie ML, Peeters A, Bulpone L, Magliano DJ, et al. Cost-effectiveness of surgically induced weight loss for the management of type 2 diabetes: modeled lifetime analysis. Diabetes Care. 2009;32:567–74.

13. Raaijmakers LC, Pouwels S, Thomassen SE, Nienhuijs SW. Quality of life and bariatric surgery: a systematic review of short- and long-term results and comparison with community norms. Eur J Clin Nutr. 2017;71:441–9.

14. Buchwald H, Avidor Y, Braunwald E, Jensen MD, Pories W, Fahrbach K, et al. Bariatric surgery: a systematic review and meta-analysis. JAMA. 2004;292:1724–37.

15. Buchwald H, Estok R, Fahrbach K, Banel D, Jensen MD, Pories WJ, et al. Weight and type 2 diabetes after bariatric surgery: systematic review and meta-analysis. Am J Med. 2009;122:248–56.

16. Heneghan HM, Meron-Eldar S, Brethauer SA, Schauer PR, Young JB. Effect of bariatric surgery on cardiovascular risk profile. Am J Cardiol. 2011;108:1499–507.

17. Romeo S, Maglio C, Burza MA, Pirazzi C, Sjöholm K, Jacobsson P, et al. Cardiovascular events after bariatric surgery in obese patients with type 2 diabetes. Diabetes Care. 2012;35:263–7.

18. Johnson BL, Blackhurst DW, Latham BB, Cull DL, Bour ES, Oliver TL, Williams B, et al. Bariatric surgery is associated with a reduction in major macrovascular and microvascular complications in moderately to severely obese patients with type 2 diabetes mellitus. J Am Coll Surg. 2013;216:545–56.

19. Busetto L, De Stefano F, Figozzo S, Segato G, De Luca M, Favretti F. Cardiovascular events after bariatric surgery on cancer incidence in obese patients in Sweden.
(Swedish obese subjects study): a prospective, controlled intervention trial. Lancet Oncol. 2009;10:653–62.

35. Zhou X, Yu J, Li L, Gloy VL, Nordmann A, Tiboni M, Li Y, Sun X. Effects of bariatric surgery on mortality, cardiovascular events, and cancer outcomes in obese patients: systematic review and meta-analysis. Obes Surg. 2016;26:2590–601.

36. Vandenbroucke JP, von Elm E, Altman DG, Gøtzsche PC, Mulrow CD, Pocock SJ, STROBE Initiative, et al. Strengthening the reporting of observational studies in epidemiology (STROBE): explanation and elaboration. PLoS Med. 2007;4:e297.

37. Sjöström L, Lindroos AK, Peltonen M, Torgerson J, Bouchard C, Carlsson B, Dahlström S, Larsson B, Narbro K, Sjöström CD, Sullivan M, Wedel H, Swedish Obese Subjects Study Scientific Group. Lifestyle, diabetes, and cardiovascular risk factors 10 years after bariatric surgery. N Engl J Med. 2004;351:2683–93.

38. Adams TD, Davidson LE, Litwin SE, Kim J, Kolotkin RL, Nanjee MN, Gutiérrez JM, Frogley SJ, Ibele AR, Brinton EA, Hopkins PN, Simper SC, Hunt SC. Weight and metabolic outcomes 12 years after gastric bypass. N Engl J Med. 2017;377:1143–55.

39. Davidson LE, Adams TD, Kim J, Jones JL, Hashibe M, Taylor D, et al. Association of patient age at gastric bypass surgery with long-term all-cause and cause-specific mortality. JAMA Surg. 2016;151:631–7.

40. Salehidoost R, Mansouri A, Arimi M, Yamin SA, Aminomroaya A. Body mass index and the all-cause mortality rate in patients with type 2 diabetes mellitus. Acta Diabetol. 2018;55:569–77.

41. Zucker I, Shoahat T, Dankner R, Chodick G. New onset diabetes in adulthood is associated with a substantial risk for mortality at all ages: a population based historical cohort study with a decade-long follow-up. Cardiovasc Diabetol. 2017;16:105.

42. Scicali R, Rosenthal D, Di Pino A, Giral P, Cluzel P, Redheuil A, Piro S, Rabuazzo AM, Purrello F, Bruckert E, Galli A. An increased waist-to-hip ratio is a key determinant of atherosclerotic burden in overweight subjects. Acta Diabetol. 2018;55:741–9.

43. Bodo A, Ceriani V, Cetta F, Lombardi F, Pontiroli AE. Bariatric surgery and prevention of cardiovascular events and mortality in morbid obesity: mechanisms of action and choice of surgery. Nutr Metab Cardiovasc Dis. 2015;25:437–43.