Probing tissue microstructure by diffusion skewness tensor imaging

Lipeng Ning1, Filip Szczepankiewicz1, Markus Nilsson2, Yogesh Rathi1 & Carl-Fredrik Westin1

Probing the cellular structure of in vivo biological tissue is a fundamental problem in biomedical imaging and medical science. This work introduces an approach for analyzing diffusion magnetic resonance imaging data acquired by the novel tensor-valued encoding technique for characterizing tissue microstructure. Our approach first uses a signal model to estimate the variance and skewness of the distribution of apparent diffusion tensors modeling the underlying tissue. Then several novel imaging indices, such as weighted microscopic anisotropy and microscopic skewness, are derived to characterize different ensembles of diffusion processes that are indistinguishable by existing techniques. The contributions of this work also include a theoretical proof that shows that, to estimate the skewness of a diffusion tensor distribution, the encoding protocol needs to include full-rank tensor diffusion encoding. This proof provides a guideline for the application of this technique. The properties of the proposed indices are illustrated using both synthetic data and in vivo data acquired from a human brain.

Noninvasive probing of cellular structure of biological tissue in vivo is a fundamental problem in medical science. Modern medical imaging modalities, such as computed tomography (CT) and magnetic resonance imaging (MRI), provide imaging data with spatial resolution at the scale of millimeter (mm). But mental disorders and other brain diseases are related to alternations in the cellular microstructure at the micrometer (µm) scale without any gross effect at the macroscopic scale. However, no in vivo imaging technique is currently available to directly probe microscopic cellular arrangements in the human brain. Diffusion MRI (dMRI) is a modality to indirectly characterizes the microscopic cellular arrangements via the diffusion trajectories of water molecules1,2.

The diffusion coefficient of water at human body temperature is approximately 3 µm²/ms. In biological tissue, e.g., human brains, the displacement of water molecules is restricted or hindered by cellular membranes. The measured diffusion coefficients using dMRI, which is usually called the apparent diffusivity, depend on the microstructure of tissue as well as the experimental parameters such as the diffusion time. Within the typical 10 to 100 ms diffusion time used in dMRI, the radius of the diffusion trajectory of a water molecule is usually in the order of several µm, similar to the size of cells. Because of microstructural heterogeneity, water molecules from different tissue components within one voxel in dMRI have different apparent diffusivity. Moreover, the diffusion trajectories can be characterized by different principal diffusion directions and different degrees of anisotropy. Thus, characterizing the statistical properties of all diffusion trajectories provide information to indirectly assess the property of tissue microstructure. Furthermore, dMRI can be performed with various diffusion encoding waveforms to sensitize the MR signal to different properties of diffusion trajectories. Then, tissue microstructure can be estimated by using a suitable analysis method for dMRI data. For example, the apparent diffusion tensor is a standard technique to characterize orientation-dependent diffusion which is related to the underlying cellular or axonal directions3. But this simplistic model only reflects the diffusion tensor of the ensemble average process with no information provided about the underlying variance and other high-order moments of the distribution of diffusivity of all water molecules that are useful to characterize the tissue heterogeneity.

In fact, this limitation is not only caused by the signal model but also the acquisition protocol. Specifically, dMRI data acquired by the standard linear-encoding sequence is not sufficient to characterize the variance of the distribution of apparent diffusion tensors4. To overcome this limitation, more advanced diffusion encoding waveforms, such as double-diffusion encoding5, q-MAS isotropic diffusion encoding7, and more general q-trajectory encoding (QTE)4,8, have been developed to characterize the microscopic anisotropy of the underlying diffusion processes. While these techniques have been used to estimate the variance of the underlying diffusion

1Brigham and Women’s Hospital, Harvard Medical School, Boston, USA. 2Lund University, Lund, Sweden. *email: lning@bwh.harvard.edu
and the corresponding standard deviation (STD) in an axial slice of the brain. The STD values were estimated using the residual bootstrap method proposed in10. The first row shows diffusion tensors with fast and slow diffusivity. As a result, the underlying tensors are of the same shape and size.

Experiment and results

Statistical indices of diffusion tensor distributions. A voxel can have tissue compartments with different cellular structures that are characterized by different apparent diffusivities at long diffusion time scale. The mixture of apparent diffusivity from these tissue compartments can be represented by the diffusion tensor distribution (DTD) function, whose statistical property indirectly characterizes the tissue heterogeneity. Figure 1 compares several statistical measures of three synthetic DTDs. The eigenvalues of the diffusion tensors in DTD1 and DTD2 are 0.1, 0.5, 0.5 m^2/ ms and 0.63, 0.233, 0.233 m^2/ ms, respectively. DTD3 has 88% of the tensors with eigenvalues being 0.63, 0.045, 0.045 m^2/ ms and 12% of isotropic tensors with eigenvalues being 1.3 m^2/ ms.

The three synthetic DTDs in Fig. 1 all have the same mean isotropic diffusion tensor. As a result, the three DTDs cannot be distinguished by any indices derived from the mean tensor, including the fractional anisotropy (FA) and the skewness (SK) see Methods. The microscopic anisotropy, μFA, has been proposed to characterize more specific information of the DTDs using the underlying variances4,7, see Methods. But the μFA of the three DTDs shown in Fig. 1 have similar values, indicating the limited sensitivity and specificity of μFA for characterizing tissue microstructure.

By using the third-order moment of the DTDs, we derived the microscopic skewness μSK and the weighted microscopic anisotropy, i.e. μFAfast and μFAslow, to characterize the statistical property of the DTDs, see Methods. In particular, μSK is sensitive to the shape of the microscopic diffusion tensors. In Fig. 1, DTD1 and DTD2 have negative and positive μSK values, respectively, which reflect the underlying oblate and prolate ellipsoids for tensor representations. On the other hand, μFAfast and μFAslow are sensitive to the microscopic anisotropy of diffusion tensors with fast and slow diffusivity. DTD3 includes isotropic tensors with relatively fast diffusivity and anisotropic tensors with relatively slow diffusivity. As a result, the underlying μFAfast is higher than μFAslow. But μFAfast and μFAslow in DTD1 and DTD2 are not different, since the underlying tensors are of the same shape and size.

Diffusion encoding waveforms. The proposed statistical indices have been applied to analyze an in vivo dataset of acquired from a human brain. The first and second rows of Fig. 2 illustrate the diffusion-encoding gradient trajectories corresponding to different shapes of b-tensors, where the waveforms along x, y and z axes are coded by different colors. In the third row of Fig. 2, the rotation directions of the waveforms are shown by the dots on the spheres, and the b-values, i.e. the trace of b-tensors, are indicated by the radius of the spheres. A total number of 513 b-tensors are used in this experiment to ensure that the proposed statistical indices are uniquely determined by the dMRI signals, see Methods.

Microstructural measures of a human brain. Figure 3 shows the estimated dMRI measures, including FA, μFA, μFAslow, μFAfast, SK, μSK and the corresponding standard deviation (STD) in an axial slice of the brain. The STD values were estimated using the residual bootstrap method proposed in10. The first row shows
the FA, the micro-FA (µFA) and their STD values. The red and yellow circles in Fig. 3a point out two white-matter regions with relative lower FA values but with uniform µFA measures compared to other white-matter regions. The STD map in Fig. 3d shows that µFA is more reliable in white-matter. The second row shows that µFA slow is higher than µFA fast especially in gray matter. It was shown in11 that extra-axonal water may have slower diffusivity than the intra-axonal water. Thus, extra-axonal water may be also related to higher µFA. Figure 3f,h show that µFA slow has lower STD values than µFA fast especially in white matter. The third row shows the SK, the micro-SK, µSK and their STD maps. The white-matter region in the yellow circle has more negative SK values than the region corresponding to the red circle, though the FA in the yellow circle is relative higher. Thus, the diffusion tensors in the yellow-circle region are more prolate and anisotropic than those from the red-circle regions. The difference in FA and SK between the two regions may be related to the number of crossing-fibers12,13, the fiber dispersion14 and/or the angle between crossing fibers15, which needs to be further examined.

The STD of SK in Fig. 3j is relatively high in white-matter, especially in regions with crossing fibers. The very high µSK values in the boundary between white/gray matter and the cerebrospinal fluid (CSF) in Fig. 3k are related to high STD values shown in Fig. 3l. It indicates that the µSK in these regions are sensitive to measurement noise, which may be related to the underlying partial volume effects, i.e. both CSF and white/gray matter are contained in the underlying voxel. Parameter uncertainty is inflated in a band toward the posterior of the brain as shown in Fig. 3b,f,h,j. This is likely due to poor fat saturation causing a mixture of fat and water signals at this location, as previously described9.

Accuracy and precision of diffusion indices. To examine the reliability of the estimated indices, the 513 b-tensors were applied to simulate dMRI signals for the three DTDs shown in Fig. 1 with additive Gaussian noise. The dMRI signals were computed by sampling \( s(B) = (e^{-\gamma B \rho} + \nu_B) + \nu \) with \( B \) representing one of the 513 b-tensors and \( \nu_B \) representing the zero-mean Gaussian noise. For a fixed standard deviation of Gaussian noise, the 513 dimensional dMRI signals were sampled 5000 times. Then, the same estimation methods as applied in the in vivo experiment were applied to estimate the µFA fast, µFA slow and µSK measures. To further examine the accuracy and precision of the indices for different DTD functions, the experimental procedure was applied to the DTD function shown in Fig. 1c with varying volume fraction of anisotropic components and the signal to noise ratio (SNR), i.e. the ratio between the mean signal and the standard deviation of noise, being fixed at SNR=30 similar to the SNR values of the in vivo data.
Figure 3. Illustration of dMRI indices and the corresponding uncertainty of an axial slice of dMRI volume from a human brain. The first row shows the FA, $\mu$FA and the corresponding standard deviation (STD). The second row illustrates the $\mu$FA, and $\mu$FA, measures, where $\mu$FA has relative higher values than $\mu$FA. The third row shows SK and $\mu$SK. SK has negative values in the white-matter region highlighted by the red and yellow circle whereas the corresponding $\mu$SK are positive, indicating that the underlying microscopic diffusion tensors have prolate shapes that are related to crossing fibers or fiber dispersion. Parameter uncertainty is inflated in a band toward the posterior of the brain. This is likely due to poor fat saturation causing a mixture of fat and water signals at this location, as previously described by 9.

Figure 4 illustrates the median values (solid lines) and 25% and 75% percentiles (the boarder lines of the shaded areas) of the estimated indices with different values of inverse signal-to-noise (1/SNR). The dashed lines indicate the underlying true values. The median values of $\mu$FA and $\mu$FA increase with higher noise levels. Moreover, the median values for $\mu$FA and $\mu$FA in Fig. 4a,b are close to each other at different noise levels. But the median values $\mu$FA and $\mu$FA in Fig. 4c are consistently far apart, indicating that the relative difference between $\mu$FA and $\mu$FA is reliable to characterize the DTDs, though these values are biased from the underlying true values. On the other hand, the median values of $\mu$SK in Fig. 4d–f decrease in accordance to increasing noise levels and all intersect with the underlying true values. The median values of $\mu$SK in Fig. 4d remain negative but the corresponding values in Fig. 4e,f gradually decrease from positive to negative values, indicating that sign of $\mu$SK is a reliable indices at low SNR values.

Figure 5 shows the estimated indices corresponding to the DTD shown in Fig. 1c with different volume fractions for the anisotropic components. Similar to Fig. 4, the solid red and blue lines show the median of the estimated values and the dashed lines show the underlying true values. The upper and lower boundaries of the shaded area correspond to 25% and 75% percentiles of the estimated measures. The true values stay within the shaded areas when the volume fraction of anisotropic components is higher than 0.4. But marked errors appear as the volume fraction decreases, indicating that partial volume effect may lead to biases in the diffusion indices. Moreover, the $\mu$SK has very high uncertainties indicating that it is challenging to be correctly estimated, which is consistent to the results of the gray matter region in Fig. 3i.
Discussion and conclusions

This work introduces a method for estimating the variance and skewness of a distribution of diffusion tensors as an approach for characterizing the microstructure of materials or biological tissue. The method is based on expanding the diffusion tensor probability distribution in its cumulant moments, and relies on the diffusion weighting with full-rank b-tensors. The skewness of diffusion tensor, diffusion skewness and high-order tensors have been investigated for tissue microstructure estimation based on standard (rank-one) b-tensor diffusion encoding (SDE). The added benefit of the proposed method with more general diffusion encoding is a novel capacity to distinguish diffusion tensors based on their skewness on the microscopic level, disentangling prolate and oblate shapes, as well as separately quantifying the microscopic anisotropy of diffusion tensors with high and low isotropic diffusivity. Taken together these features can add additional information on tissue microstructure, expanding on lower order methodology like DKI and QTI, which have been used to characterize prostate cancer and intracranial tumors. We note that the proposed model ignores the intra-compartmental...
non-Gaussian diffusion\textsuperscript{24,25} and time-dependent diffusivity\textsuperscript{26,27}, which can be further integrated to the model to improve the accuracy for microstructure estimation. We do not expect these effects are significant in the healthy brain especially in white matter\textsuperscript{28,29}. Moreover, the proposed indices based on cumulant moments, similar to DKI and QTI, do not directly provide specific information about tissue microstructure such as intra-axonal/cellular volume fractions, though they are sensitive to microstructural changes. From studies using cumulant expansions up to the second order we know that there exists a tradeoff between trueness and precision, where the use of higher b-values typically increase the precision while being detrimental to the trueness\textsuperscript{30}. Nevertheless, metrics provided by such approaches can be useful in radiology for example to improve glioma grading\textsuperscript{31}. The trueness of the parameters proposed here certainly also depends on the specifics of the imaging protocol. Further studies are needed to carefully balance the trueness and precision in this context. We emphasize the need for further investigations of tissues of pathological conditions where this may provide relevant information. We speculate that it may resolve the equality of prolate and oblate solutions and being able to robustly detect the presence of planar structures, such as those conjectured to exist in white matter sheets\textsuperscript{32}. Currently, a proof-of-concept is illustrated in a healthy human brain in vivo. Indices derived from the model will be validated using liquid crystal phantoms\textsuperscript{32} and biological tissue in future works.

In summary, this work has introduced an approach for modeling and analyzing diffusion MRI data acquired by the novel sampling scheme with a wide range of b-tensor shapes. The main contributions include several statistical indices based on the skewness tensor of DTD functions and a mathematical proof that full-rank b-tensors are needed to uniquely determine the skewness tensor.

**Methods**

**Modeling of diffusion MRI signals.** In a diffusion experiment, on the application of a time-varying magnetic gradient field, $g(t) \in \mathbb{R}^3$ for $t \in [0, \tau]$, the phase change of a rotating spin due to diffusive motion of water molecule during this time window is given by\textsuperscript{33}:

$$\phi_x(t) = \gamma \int_0^\tau r_x(t) g(t) dt = -\int_0^\tau \nu_x(t)^T q(t) dt,$$

where $\gamma$ is the gyromagnetic ratio, $r_x(t)$ denotes the trajectory of the displacement of a particle starting from $x$, $\nu_x(t)$ denotes the corresponding velocity process, which is formally defined as $\nu_x(t) = \dot{r}_x(t)$, $q(t) = \gamma \int_0^\tau g(s) ds$ is the q-trajectory and $\dot{T}$ denotes the transpose. If the diffusion trajectory is assumed to be a three-dimensional Wiener process with constant diffusivity $D$ during the period of diffusion time, then $\phi_x(t)$ follows a zero-mean Gaussian distribution with the covariance equal to

$$\mathbf{C}(\phi_x^2(t)) = \gamma^2 \int_0^\tau q(t)^T D q(t) dt = \text{trace}(BD) \triangleq B : D,$$

where $B = \int_0^\tau q(t)^{02} dt$ is the b-tensor. Then, the expected value of the corresponding MR signal is equal to $\mathbf{E}(\phi_x(t)) = e^{-BD}$.

The diffusion-weighted MR signal after being normalized by the non-diffusion-weighted signal, i.e. the baseline, is equal to the ensemble average of signals from all water molecules given by:

$$s(B) = \langle e^{\phi(t)} \rangle_p$$

where $\langle \cdot \rangle_p$ denotes the ensemble average with respect to the phase distribution function $p(\phi)$. At long diffusion time scale when the travel distances of water molecules much longer than characteristic cellular sizes, the molecular diffusion processes can be approximated by Gaussian processes with location dependent diffusivity\textsuperscript{34}. At long diffusion time, if the mixture of apparent diffusion tensors from different cellular compartments within a voxel follows the diffusion tensor distribution (DTD) function $\rho(D)$, then the diffusion-weighted signal is equal to\textsuperscript{34,35}:

$$s(B) = \int e^{-BD} \rho(D) dD \triangleq \langle e^{-BD} \rangle_\rho.$$

A similar signal model based on scalar-valued diffusivity was investigated in\textsuperscript{36}.

**The cumulant expansion.** The standard dMRI approaches probe the diffusion process using a linear diffusion gradient waveform which corresponds to a b-tensor of rank one. In this case, estimating the underlying DTD $\rho(D)$ using an inverse Laplace transform is a highly ill-posed problem. However, advanced QTE sequences are able to probe $\rho(D)$ using any b-tensor up to rank 3, providing information that cannot be measured by SDE\textsuperscript{3}. The ensemble average signal $\langle e^{-BD} \rangle_\rho$ can be approximated using the following cumulant expansion:

$$\log s(B) \approx -B : \langle D \rangle_\rho + \frac{1}{2} B^{02} : C - \frac{1}{6} B^{03} : \mathcal{S},$$

where $\langle D \rangle_\rho$, $C$, $\mathcal{S}$ denote the first three cumulants and are given by: $C = \langle (D - \langle D \rangle_\rho)^{02} \rangle_\rho$ and $\mathcal{S} = \langle (D - \langle D \rangle_\rho)^{03} \rangle_\rho$. The notation $\otimes$ denotes the tensor product and $\cdot$ defines the standard inner product between matrices. For example, $B^{03} : \mathcal{S} = [B^{03}]_{ijkl} [\mathcal{S}]_{klm}^i$ using Einstein summation convention. The tensors $\langle D \rangle_\rho$, $C$ and $\mathcal{S}$ have 6, 21 and 56 independent variables, respectively. These 83 variables can be estimated by solving the following linear system of equations:
where \( \overrightarrow{B} \), \( \overrightarrow{D} \), and \( \overrightarrow{\mathcal{C}} \) are 6-dimensional column vector representations for \( D \) and \( B \) using the Mandel-Voigt notations such that \( \overrightarrow{B} \cdot \overrightarrow{D} = B : \overrightarrow{D} \). Similarly, \( \overrightarrow{\mathcal{C}} \) and \( \overrightarrow{\mathcal{F}} \) are 21- and 56-dimensional column-vector representations for \( \overrightarrow{C} \) and \( \overrightarrow{\mathcal{F}} \), respectively. Clearly, the number of measurements, \( m \), needs to be at least 83 to ensure that the measurement matrix (\( m \times 83 \)) is non-singular. The 513 b-tensors used in the experiment provides a full-rank measurement matrix of size 513 \( \times 83 \). Next we show that non-singular b-tensors are needed to ensure that the measurement matrix is of full rank.

**On full-rank b-tensors.** In linear encoding sequences, the gradient sequences have varying magnitudes along a fixed direction. In this case, the corresponding b-tensors have rank one. But the corresponding 21-dimensional vector \( \overrightarrow{B^{(2)}} \) only spans a 15-dimensional subspace. Consequently, the variable \( \overrightarrow{C} \) in the linear system (3) cannot be uniquely determined using standard SDE measurements. However, the vectors \( \overrightarrow{B^{(2)}} \) corresponding to rank-2 b-tensors given by planar encoding sequences\(^5^,\)\(^6^,\)\(^7^,\)\(^8\) or rank-3 tensors general QTE sequences\(^9\) are able to generate full-rank 21-dimensional space, providing a unique solution to \( \overrightarrow{C} \). But b-tensors of rank two still have zero determinant, i.e.

\[
\det(\overrightarrow{B}) = b_{11} b_{22} b_{33} + 2 b_{12} b_{23} b_{13} - b_{13}^2 b_{22} - b_{12}^2 b_{33} - b_{11} b_{23} = 0,
\]

which implies that the corresponding elements in \( \overrightarrow{B^{(2)}} \) are linearly dependent. Therefore using only b-tensors of rank one or rank two is not sufficient to generate full-rank measurement matrices in Eq. (3). This shows that a full-rank b-tensor is needed to ensure a unique solution to Eq. (3). The importance of rank-2 b-tensors in estimating the covariance of DTDs was considered in\(^5^,\)\(^8\). The above analysis further extends the theory to show the importance of full-rank b-tensors in estimating the skewness of DTDs.

**Microscopic anisotropy and skewness.** The estimated moments (or cumulants) can be used to derive scalar indices to characterize the underlying DTDs. In particular, the standard diffusion tensor imaging (DTI) technique uses the mean diffusion tensor \( \overrightarrow{D} \) to compute several classical indices including the mean diffusivity, \( \text{MD} = \frac{1}{3} \text{tr}(\overrightarrow{D}) \), and the fractional anisotropy \( \text{FA} \). But these measures provide limited information about the underlying DTD. For example, the mean diffusion tensor of the three DTDs illustrated in Fig. 1 is proportional to the identity matrix with \( \text{FA} = 0 \), though the underlying DTD functions are different. One can look at the covariance tensor \( \overrightarrow{C} \), and compute the micro-anisotropy \( \mu \text{FA} \) of the DTDs given by:

\[
\mu \text{FA}(D) = \frac{3}{2} \frac{1}{\overline{\overrightarrow{D}^{(2)}}} \text{Cov}(\overrightarrow{D})_\rho = \frac{3}{2} \left( \overrightarrow{D}^{(2)} \cdot \mathbb{E}_{\text{shear}} \right)_\rho,
\]

where \( \lambda_i(D) \) denotes the \( i \)-th eigenvalue of \( D \), \( \overline{\lambda}(D) = \frac{1}{3} \sum_{i=1}^{3} \lambda_i(D) \), \( \mathbb{E}_{\text{iso}} = \frac{1}{3} I_{6 \times 6} \) with \( I_{6 \times 6} \) being the identify matrix of size \( 6 \times 6 \) and \( \mathbb{E}_{\text{shear}} \) is equal to\(^4\)

\[
\mathbb{E}_{\text{shear}} = \frac{1}{9} \begin{bmatrix}
2 & -1 & -1 & 0 & 0 & 0 \\
-1 & 2 & -1 & 0 & 0 & 0 \\
-1 & -1 & 2 & 0 & 0 & 0 \\
0 & 0 & 0 & 3 & 0 & 0 \\
0 & 0 & 0 & 0 & 3 & 0 \\
0 & 0 & 0 & 0 & 0 & 3
\end{bmatrix}.
\]

However, in many scenarios, even the variance tensor is not enough to distinguish between the DTDs. For example, the three DTD functions in Fig. 1 have the same \( \mu \text{FA} \) measure though the underlying microscopic diffusion tensors are different. In this case, the third-order moments provide useful information to distinguish the underlying DTD functions. To this end, we propose two sets of indices, \( \mu \text{FA}_{\text{fast}}, \mu \text{FA}_{\text{slow}} \) and \( \mu \text{SK}, \mu \text{SK}_{\text{fast}}, \mu \text{SK}_{\text{slow}} \), to characterize the DTDs using the third-order moments.

To introduce the first approach, we introduce a family of DTD functions defined by:

\[
\rho_f(D) = \frac{f(D)}{\langle f(D) \rangle_\rho} \rho(D),
\]

where \( f(D) \) is a given positive scalar-valued function of the diffusion tensor \( D \). This function is chosen to scale the probability density so that the statistical property of the filtered DTD function \( \rho_f(D) \) provides more specific information about the underlying distribution. The moments of the filtered DTD function \( \rho_f(D) \) can be computed using:
\[
(D^{\otimes k})_j = \int D^{\otimes k} \rho(D) dD = \frac{\langle D^{\otimes k} f(D) \rangle_\rho}{\langle f(D) \rangle_\rho}.
\]

If \(f(D)\) is a polynomial function of \(D\), then both the numerator and the denominator of the above equation are linear combinations of the moments of \(\rho(D)\). In particular, we consider the following two linear functions:

\[ f_{\text{fast}}(D) = \text{trace}(D), \]

\[ f_{\text{slow}}(D) = \hat{d} - \text{trace}(D), \]

where \(\hat{d}\) is a positive scalar such that \(f_{\text{slow}}(D) > 0\) for all feasible diffusion tensors. The functions \(f_{\text{fast}}(D)\) and \(f_{\text{slow}}(D)\) selectively weigh the density \(\rho(D)\) with relatively high and low trace values. We denote the corresponding filtered DTD function by \(\rho_{\text{fast}}\) and \(\rho_{\text{slow}}\), respectively. The difference between the moments of \(\rho_{\text{fast}}\) and \(\rho_{\text{slow}}\) reflects the distribution of \(\rho(D)\) over tensors with different diffusivities.

From (6), the first and second order moments of \(\rho_{\text{fast}}(D)\) is equal to:

\[ \langle D \rangle_{\rho_{\text{fast}}} = \frac{\langle \text{trace}(D) \rangle_{\rho}}{\langle \text{trace}(D) \rangle_{\rho}}, \]

\[ \langle D^{\otimes 2} \rangle_{\rho_{\text{fast}}} = \frac{\langle \text{trace}(D) D^{\otimes 2} \rangle_{\rho}}{\langle \text{trace}(D) \rangle_{\rho}}, \]

which can be computed using the estimated moments of \(\rho(D)\). Similarly, the first and second order moments of \(\rho_{\text{slow}}(D)\) are equal to:

\[ \langle D \rangle_{\rho_{\text{slow}}} = \frac{\hat{d}(D)_{\rho} - \langle \text{trace}(D) \rangle_{\rho}}{\hat{d} - \langle \text{trace}(D) \rangle_{\rho}}, \]

\[ \langle D^{\otimes 2} \rangle_{\rho_{\text{slow}}} = \frac{\hat{d}(D^{\otimes 2})_{\rho} - \langle \text{trace}(D) D^{\otimes 2} \rangle_{\rho}}{\hat{d} - \langle \text{trace}(D) \rangle_{\rho}}. \]

These equations can be used to estimate specific dMRI measures using the estimated moments for \(\rho_{\text{fast}}(D)\) and \(\rho_{\text{slow}}(D)\). For the DTD functions shown in Fig. 1, the \(\mu FA\) measures for \(\rho_{\text{fast}}(D)\) and \(\rho_{\text{slow}}(D)\), denoted by \(\mu FA_{\text{fast}}\) and \(\mu FA_{\text{slow}}\), corresponding to DTD1 and DTD2 have the same values since the underlying diffusion tensors have the same trace. For DTD3, \(\mu FA_{\text{fast}}\) is much lower than \(\mu FA_{\text{slow}}\) since the underlying fast diffusion tensors are isotropic whereas the slow-diffusion components are anisotropic.

To introduce the microscopic skewness, we first extend the skewness measure for probability distribution functions to define the following measure of asymmetry of the distribution of eigenvalues of the mean diffusion tensor \(\langle D \rangle\):

\[
\text{SK}(\langle D \rangle) = \frac{1}{\frac{1}{3} \sum_{i=1}^{3} \langle \lambda_{i}(\langle D \rangle) - \bar{\lambda}(\langle D \rangle) \rangle^3}{\langle (\langle D \rangle)^{\otimes 3} : \hat{\varepsilon}_{\text{shear}} \rangle_{\rho} \langle \hat{\varepsilon}_{\text{shear}} \rangle_{\rho}^{3/2}},
\]

where \(\langle D \rangle\) is assumed to be anisotropic so that the denominator is assumed nonzero. In above, \(\hat{\varepsilon}_{\text{shear}}\) is a fully symmetric three-dimensional tensor, i.e. the value of \([\hat{\varepsilon}_{\text{shear}}]_{ijk}\) does not change by permuting the order of \(i,j,k\), of size \(6 \times 6 \times 6\) with the following non-zero entries:

\[ [\hat{\varepsilon}_{\text{shear}}]_{iii} = \frac{2}{27} \text{ if } i \in \{1, 2, 3\}, \]

\[ [\hat{\varepsilon}_{\text{shear}}]_{ijj} = -\frac{1}{27} \text{ if } i,j \in \{1, 2, 3\} \text{ and } i \neq j, \]

\[ [\hat{\varepsilon}_{\text{shear}}]_{iij} = \frac{1}{18} \text{ if } i \in \{4, 5, 6\} \text{ and } j \in \bar{\mathcal{J}}, \]

\[ [\hat{\varepsilon}_{\text{shear}}]_{iji} = -\frac{1}{18} \text{ if } i \in \{4, 5, 6\} \text{ and } j \in \{1, 2, 3\} \backslash \bar{\mathcal{J}}, \]

\[ [\hat{\varepsilon}_{\text{shear}}]_{ijk} = \frac{\sqrt{2}}{12} \text{ if } i,j,k \in \{4, 5, 6\}, \]

where \(\bar{\mathcal{J}} = \{1, 2\}, \bar{\mathcal{J}}_2 = \{1, 3\}, \bar{\mathcal{J}}_6 = \{2, 3\}\). The skewness measure SK is able to distinguish diffusion tensors that cannot be separated by FA values. For example, the diffusion tensors represented by oblate tensors have a negative SK while prolate tensors have positive SK. Following the definition of \(\mu FA\), we define the following microscopic skewness measure:

\[
\mu \text{SK}(\rho) = \frac{\langle (\langle D \rangle^{\otimes 3} : \hat{\varepsilon}_{\text{shear}} \rangle_{\rho} \langle \hat{\varepsilon}_{\text{shear}} \rangle_{\rho}^{3/2} \rangle_{\rho}}{\langle D^{\otimes 2} \rangle_{\rho}^{3/2}}.
\]
which characterizes the microscopic skewness of the diffusion tensors. To enhance robustness with respect to measurement noise, the $\langle D_{\text{52}} \rangle_p : E_{\text{shear}}$ term in the denominator on the right hand side of the above equation is replaced by $\langle D_{\text{52}} \rangle_p : E_{\text{shear}} + \epsilon$ with $\epsilon = 0.03 \mu m^2/s^2$, which does not change the sign of $\mu SK(\rho)$. In particular, the $\mu SK(\rho)$ for DTD$_1$ in Fig. 1 is negative while for DTD$_2$ and DTD$_3$ it is positive.

**Experimental parameters of the in vivo dataset.** The “free waveform” prototype pulse sequence based on a diffusion-weighted spin-echo sequence with EPI readout was used to acquire MRI data from an adult male subject. Imaging parameters were: TE = 100 ms, TR = 2800 ms, field of view = $220 \times 220 \times 75$ mm$^3$, partial Fourier factor = 6/8 and in-plane acceleration factor = 2. The acquired images were corrected for motion and eddy current distortion using an extrapolation-based approach capable of correcting diffusion weighted volumes acquired with high b-values.$^{39,40}$

A total number of 513 image volumes were acquired using different QTE trajectories tailored to the MRI system by numerical optimization$^{41}$ with compensation for concomitant gradient effects$^{42}$. The QTE waveforms were determined by two parameters $b_3$ and $b_6$, which characterize the shape of b-tensors.$^{9,43}$ The orientation of b-tensors was determined via rotation of the waveforms and b-values were changed by varying the magnitude of waveforms$^{44}$. The diffusion encoding waveforms in the first two columns provide spherical and oblate b-tensors with full-rank, which are required to estimate the skewness of the underlying DTD. The 513 b-tensors in this experiment provide a full-rank measurement matrix so that the proposed statistical indices are uniquely determined by the dMRI signals.

Received: 30 June 2020; Accepted: 30 November 2020
Published online: 08 January 2021

**References**

1. Stejskal, E. & Tanner, J. Spin diffusion measurements: spin echoes in the presence of a time-dependent field gradient. *J. Chem. Phys.* 42, 288–292 (1965).
2. Stejskal, E. & Tanner, J. Measurement of translational displacement probabilities by NMR: an indicator of compartmentation. *J. Chem. Phys.* 14, 435–444 (1990).
3. Basser, P., Mattiello, J. & LeBihan, D. Estimation of the effective self-diffusion tensor from the NMR spin echo. *J. Magn. Reson. Ser.* B 103, 247–254 (1994).
4. Westin, C.-F. et al. Q-space trajectory imaging for multidimensional diffusion MRI of the human brain. *NeuroImage* 135, 345–362 (2016).
5. Mitra, P. P. Multiple wave-vector extensions of the NMR pulsed-field-gradient spin-echo diffusion measurement. *Phys. Rev. B* 51, 15074–15078 (1995).
6. Callaghan, P. T. & Furó, I. Diffusion-diffusion correlation and exchange as a signature for local order and dynamics. *J. Chem. Phys.* 120, 4032–4038 (2004).
7. Lasič, S., Szczepankiewicz, F., Eriksson, S., Nilsson, M. & Topgaard, D. Microanisotropy imaging: quantification of microscopic diffusion anisotropy and orientational order parameter by diffusion MRI with magic-angle spinning of the q-vector. *Front. Phys.* 2, 11 (2014).
8. Westin, C.-F. et al. Measurement tensors in diffusion MRI: generalizing the concept of diffusion encoding. *Med. Image Comput. Comput. Assist. Interv.* 17, 209–216 (2014).
9. Szczepankiewicz, F., Sjödén, J., Ställberg, E., Lätt, J. & Nilsson, M. Tensor-valued diffusion encoding for diffusion variance decomposition (DIVIDE): technical feasibility in clinical MRI systems. *PLoS ONE* 14, 1–20 (2019).
10. Chung, S., Lu, Y. & Henry, R. G. Comparison of bootstrap approaches for estimation of uncertainty of DTI parameters. *NeuroImage* 33, 531–541 (2006).
11. Dhital, B., Reisert, M., Kellner, E. & Kiselev, V. G. Intra-axonal diffusivity in brain white matter. *NeuroImage* 189, 543–550 (2019).
12. Jellison, B. J. et al. Diffusion tensor imaging of cerebral white matter: a pictorial review of physics, fiber tract anatomy, and tumor imaging patterns. *Am. J. Neuroradiol.* 25, 356–369 (2004).
13. Jeurissen, B., Leemans, A., Tournier, J.-D., Jones, D. K. & Sijbers, J. Investigating the prevalence of complex fiber configurations in white matter tissue with diffusion magnetic resonance imaging. *Hum. Brain Map.* 34, 2747–2766 (2013).
14. Ronen, I. et al. Microstructural organization of axons in the human corpus callosum quantified by diffusion-weighted magnetic resonance spectroscopy of N-acetylaspartate and post-mortem histology. *Brain Struct. Funct.* 219, 1173–1185 (2014).
15. Tax, M. W. et al. Quantifying the brain's sheet structure with normalized convolution. *Med. Image Anal.* 39, 162–177 (2017).
16. Ennis, D. B. & Kindlmann, G. Orthogonal tensor invariants and the analysis of diffusion tensor magnetic resonance images. *Magn. Reson. Med.* 55, 136–146 (2006).
17. Zhang, X., Ling, C., Qi, L. & Wu, E. The measure of diffusion skewness and kurtosis in magnetic resonance imaging. *Pac. J. Optim.* 6, 391–404 (2010).
18. Özarslan, E. & Mareci, T. H. Generalized diffusion tensor imaging and analytical relationships between diffusion tensor imaging and high angular resolution diffusion imaging. *Magn. Reson. Med.* 50, 955–965 (2003).
19. Liu, C., Bammer, R., B. A. & Moseley, M. E. Characterizing non-gaussian diffusion by using generalized diffusion tensors. *Magn. Reson. Med.* 51, 924–937 (2004).
20. Liu, C., Mang, S. C. & Moseley, M. E. In vivo generalized diffusion tensor imaging (GDTI) using higher-order tensors (HOT). *Magn. Reson. Med.* 63, 243–252 (2010).
21. Jensen, J. H., Helpern, J. A., Ramani, A., Lu, H. & Kaczyński, K. Diffusional kurtosis imaging: the quantification of non-gaussian water diffusion by means of magnetic resonance imaging. *Magn. Reson. Med.* 53, 1432–1440 (2005).
22. Roetkehe, M. et al. Evaluation of diffusion kurtosis imaging versus standard diffusion imaging for detection and grading of peripheral zone prostate cancer. *Invest. Radiol.* 50, 483–489 (2015).
23. Nilsson, M. et al. Tensor-valued diffusion MRI in under 3 minutes: an initial survey of microscopic anisotropy and tissue heterogeneity in intracranial tumors. *Magn. Reson. Med.* 83, 608–620 (2020).
24. Neto Henriques, R., Jespersen, S. N. & Shemesh, N. Microscopic anisotropy misestimation in spherical-mean single diffusion encoding MRI. *Magn. Reson. Med.* 81, 3245–3261 (2019).
25. Neto Henriques, R., Jespersen, S. N. & Shemesh, N. Correlation tensor magnetic resonance imaging. *NeuroImage* 211, 116605 (2020).
26. Jespersen, S. N., Olesen, J. L., Ianus, A. & Shemesh, N. Effects of nongaussian diffusion on isotropic diffusion measurements: an ex-vivo microimaging and simulation study. *J. Magn. Reson.* 300, 84–94 (2019).
27. Ning, L., Setsompop, K., Westin, C.-F. & Rathi, Y. New insights about time-varying diffusivity and its estimation from diffusion MRI. *Magn. Reson. Med.* 78, 763–774 (2017).

28. Szczepankiewicz, F. et al. Is spherical diffusion encoding rotation invariant? An investigation of diffusion time-dependence in the healthy brain. In ISMRM 27th Annual Meeting and Exhibition, 0223 (2019).

29. Szczepankiewicz, F., Westin, C.-F. & Nilsson, M. Gradient waveform design for tensor-valued encoding in diffusion MRI. *ArXiv e-prints* (2020). arXiv:2007.07631.

30. Chuhutin, A., Hansen, B. & Jespersen, S. N. Precision and accuracy of diffusion kurtosis estimation and the influence of b-value selection. *NMR Biomed.* 30, e3777 (2017).

31. Delgado, M. et al. Chaski, a novel Drosophila lactate/pyruvate transporter required in glia cells for survival under nutritional stress. *Sci. Rep.* 8, 1186 (2018).

32. Nilsson, M. et al. Liquid crystal phantom for validation of microscopic diffusion anisotropy measurements on clinical MRI systems. *Magn. Reson. Med.* 79, 1817–1828 (2017).

33. Callaghan, P. T. *Principles of Nuclear Magnetic Resonance Microscopy* (Clarendon Press, Oxford, 1994).

34. Yablonskiy, D. A. & Sukstanskii, A. L. Theoretical models of the diffusion weighted MR signal. *NMR Biomed.* 23, 661–681 (2010).

35. Topgaard, D. Diffusion tensor distribution imaging. *NMR Biomed.* 32, e4066 (2019).

36. Yablonskiy, D. A., Brethorst, G. L. & Ackerman, J. J. Statistical model for diffusion attenuated mr signal. *Magn. Reson. Med.* 50, 664–669 (2003).

37. Jensen, J. H., Hui, E. S. & Helpern, J. A. Double-pulsed diffusional kurtosis imaging. *NMR Biomed.* 27, 363–370 (2014).

38. Shemesh, N. et al. Conventions and nomenclature for double diffusion encoding NMR and MRI. *Magn. Reson. Med.* 75, 82–87 (2016).

39. Nilsson, M., Szczepankiewicz, F., van Westen, D. & Hansson, O. Extrapolation-based references improve motion and eddy-current correction of high b-value dwi data: application in parkinson’s disease dementia. *PLoS ONE* 10, 1–22 (2015).

40. Klein, S., Staring, M., Murphy, K., Viergever, M. A. & Pluim, J. P. W. Elastix: a toolbox for intensity-based medical image registration. *IEEE Trans. Med. Imaging* 29, 196–205 (2010).

41. Sjölund, J. et al. Constrained optimization of gradient waveforms for generalized diffusion encoding. *J. Magn. Reson.* 261, 157–168 (2015).

42. Szczepankiewicz, F., Westin, C. & Nilsson, M. Maxwell-compensated design of asymmetric gradient waveforms for tensor-valued diffusion encoding. *Magn. Reson. Med.* 82, 1424–1437 (2019).

43. Topgaard, D. Multidimensional diffusion MRI. *J. Magn. Reson.* 275, 98–113 (2017).

44. Szczepankiewicz, F., Hoge, S. & Westin, C.-F. Linear, planar and spherical tensor-valued diffusion MRI data by free waveform encoding in healthy brain, water, oil and liquid crystals. *Data Brief* 25, 104208 (2019).

Acknowledgements
The authors are grateful for the support by NIH under grants P41EB015902, R01MH074794, R01MH111917, R01MH119222, R01MH116173, R21MH115280, R21MH116352, K01MH117346.

Author contributions
L.N. wrote the manuscript text. L.N. and F. S. prepared figures. M.N. provided a part of the code for simulations. F.S. and C.-F. W. conducted the in vivo experiment. F.S., Y.R and C.-F. W. revised the manuscript. All authors reviewed the manuscript.

Competing interests
The authors declare no competing interests.

Additional information
Correspondence and requests for materials should be addressed to L.N.

Reprints and permissions information is available at www.nature.com/reprints.

Publisher’s note
Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.

Open Access This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons licence, and indicate if changes were made. The images or other third party material in this article are included in the article’s Creative Commons licence, unless indicated otherwise in a credit line to the material. If material is not included in the article’s Creative Commons licence and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this licence, visit http://creativecommons.org/licenses/by/4.0/.

© The Author(s) 2021