I. INTRODUCTION

The advent of nonlinear dynamics, originally a discipline in pure mathematics, in the physical sciences at about three decades ago triggered an avalanche of new developments, which still to date has not elapsed at all. In particular, notions like chaos and sensitive dependence on initial conditions put new emphasis on the study how small perturbations change the dynamics of physical systems. That gave new stimulus to topics like the classical Kramers problem of thermal activation [1], which originated in physical chemistry, and which nowadays is considered as one of the early milestones in the study of complex behaviour subjected to noise [2]. In the wake of such renewed interest various quite counterintuitive discoveries have been made, most notably the ability of weak noise to increase the regularity of a signal. Depending on the underlying mechanism different terms have been coined for these phenomena. Stochastic resonance refers to an increase of a signal-to-noise ratio which follows the spirit of the original Kramers problem subjected to periodic driving [3]. In the setting of a transport problem the concept of stochastic resonance gave rise to new mechanisms for nonequilibrium transport [4]. The term coherence resonance refers to the nonmonotonic dependence of the coherence or regularity upon noise intensity in an autonomous excitable system without external periodic driving, which leads to an optimum coherence of noise-induced oscillations at non-zero noise intensity [5–7]. Originally the term coherence resonance has been restricted to excitable systems where mainly numerical but also some analytical results were obtained [22–27]. Delayed coupled Hopf normal forms have been studied in [28] with numerical and analytical methods using linear response theory. Further progress in analytical terms is hampered by the observation that stochastic delay differential equations

increasing noise the excitation across threshold occurs more frequently, and thus the interspike intervals become more regular. On the other hand, with increasing noise the deterministic spiking dynamics becomes smeared out, which counteracts the former effect. Such behaviour has been shown theoretically or experimentally in a variety of excitable systems, like lasers with saturable absorber [8], optical feedback [9–11], and optical injection [12], semiconductor superlattices [13], or neural systems [6, 7]. The mechanism of coherence resonance close to bifurcations has been investigated numerically, where characteristic signatures of noisy precursors of a dynamical instability appear in the power spectrum [14, 15].

Here we want to focus on coherence resonance in non-excitable systems where the underlying mechanism is associated with a subcritical Hopf bifurcation and a saddle-node bifurcation of a pair of (stable and unstable) periodic orbits [16, 17]. The fundamental aspects and the mechanism of such a novel coherence resonance scenario are elaborated in [18]. Unlike coherence resonance in excitable systems, which is linked to two competing timescales, the main mechanism of this novel resonance is related to a transient phase space structure - a ghost of a saddle-node bifurcation of periodic orbits. In this respect the coherence resonance mechanism has something in common with type I intermittent dynamics. But contrary to the latter, the deterministic dynamics in our case does not provide a reinjection mechanism which here solely depends on the stochastic forcing and thus gives rise to the resonance characteristics. The essentially new element we want to add here is the investigation of impact of time-delayed feedback. Dynamics with time delay, even though well studied in engineering science [19], has become only recently a rapidly developing subject in applied nonlinear dynamical systems theory (see e.g. [20, 21] for recent accounts). Previous studies of coherence resonance in the presence of time delay were restricted to excitable systems where mainly numerical but also some analytical results were obtained [22, 27]. Delay-coupled Hopf normal forms have been studied in [28] with numerical and analytical methods using linear response theory. Further progress in analytical terms is hampered by the observation that stochastic delay differential equa-
tions do not define a priori a Markov process, so that all the powerful tools like Fokker-Planck equations are not available. There have been attempts to reduce stochastic delay differential systems to Fokker-Planck equations, but these approaches are either limited to linear systems, or they involve some quite severe approximations which are difficult to control \cite{29}. Hence, we also consider the problem we are dealing with, the impact of time delay on coherence resonance, as opportunity to develop and to present some of the analytical techniques that can be applied for the analysis of general nonlinear stochastic time delay systems. While these techniques are variants of well established approaches in dynamical systems theory, their application to nonlinear stochastic time delay dynamics is certainly not common knowledge, and will be presented here in a coherent form.

We focus on a paradigmatic model, a complex Hopf normal form subjected to noise and time-delayed feedback

\[
\dot{z}(t) = (\lambda + i\omega + s|z(t)|^2 - |z(t)|^4)z(t) - K(z(t) - z(t - \tau)) + \sqrt{2D\zeta(t)},
\]

where \(z \in \mathbb{C}, \lambda \in \mathbb{R}\) and \(s \in \mathbb{R}\) are the deterministic bifurcation parameters, \(\omega\) denotes the intrinsic frequency of the system, \(D \geq 0\) is the noise intensity, \(K \in \mathbb{R}\) is the strength of the feedback, and \(\tau\) is the delay time. \(\zeta(t)\) denotes normalised complex valued isotropic Gaussian white noise \((\zeta(t)\zeta^{*}(t')) = 2\delta(t-t'),\) and we focus solely on the subcritical case \(s > 0\). Then the system without time delay and without noise, \(K = 0\) and \(D = 0\), shows the classic scenario of a subcritical Hopf bifurcation with the creation of an unstable periodic orbit at \(\lambda = 0\), bistability between the stable trivial fixed point and a stable periodic orbit in the range \(-s^2/4 < \lambda < 0\), and the collision of stable and unstable periodic orbits at \(\lambda = -s^2/4,\) see figure 1. The parameter region \(\lambda < -s^2/4\) is the region of interest from the point of view of coherence resonance. The time-delayed feedback implemented in the model, eq. (1), is similar to time-delayed feedback control of deterministic orbits \cite{30}, which is known to be able to enhance and stabilise oscillatory behaviour (see e.g. \cite{31} for a comprehensive account on this and related subjects). Hence, a feedback of such a type is an obvious choice to investigate the impact of time delay on coherence resonance.

We are now going to address two issues separately. In section II we will discuss in detail the properties of the stationary probability distribution of eq. (1), in particular with regards to its dependence on the time delay. We will develop an approach along the lines of a centre manifold reduction, which allows us to derive closed analytical expressions for the stationary probability distribution. Section III is devoted to the investigation of correlation properties of eq. (1), shifting the focus to the core subject of resonance properties of the underlying dynamics. Here as well we will put the emphasis on a scheme, which will allow us to compute analytically the correlation function and the corresponding power spectrum with high accuracy. Moreover, we will introduce a novel quantity - ghost weight - to quantify coherence resonance that occurs, though not in a pronounced form, even outside the regime of bimodal stationary probability distributions. Finally, in section IV we conclude with remarks how the ideas developed in the previous sections can be merged for the benefit to understand and characterise coherence resonance in stochastic time delay systems. Some of the analytical techniques - the bifurcation analysis of the deterministic model and the centre manifold reduction of the stochastic equations - are summarised in two appendices, for the convenience of the reader.

II. STATIONARY BEHAVIOUR AND STOCHASTIC BIFURCATIONS

It is not common practise to look at resonance phenomena in stochastic systems from a traditional dynamical systems point of view, i.e., in terms of bifurcations, instabilities and topological changes in phase space. In fact, the plain study of stationary properties of stochastic systems can already signal the occurrence of resonance phenomena and can contribute to uncover new mechanisms. Here, we are going to focus on the evaluation of the stationary probability distribution of the stochastic delay system, eq. (1). Before we investigate the impact of time delay let us first briefly recall how this simple bifurcation structure is affected by noise, i.e., let us consider first the case \(K = 0\) and \(D > 0\).

There exist various nonequivalent notions of bifurcations in random systems. The approach used predominately in the mathematical context, see, e.g., \cite{32}, puts these notions on a rigorous basis. There are two

\[\lambda < -s^2/4\]
types of stochastic bifurcations: phenomenological bifurcations (P-bifurcations) and dynamical bifurcations (D-bifurcations). While the first type denotes a qualitative change of the shape of the distribution, the latter describes a change of stability for the trajectories. But such advanced theoretical concepts are often difficult to investigate and almost impossible to apply in a plain experimental context. The complementary notion of noise-induced phase transitions has less appeal from the rigorous point of view, but is straightforward to evaluate, see e.g. [33]. Here one solely focuses on the shape of the stationary distribution and tracks, in particular, the number of maxima of such a distribution to detect a stochastic P-bifurcation. While such an approach depends on the coordinate system it has the nice feature that one can easily evaluate transitions even if just an experimental time series is available. Here we follow the latter approach, which has been used in investigations of coherence resonance, see for instance [18].

If we consider the stochastic differential equation for $K = 0$, the stationary probability distribution can be written down if we solve the corresponding stationary Fokker-Planck equation [24]. In polar coordinates $z = r \exp(i \phi)$ the spherically symmetric probability distribution reads

$$P(r) = N r \exp \left( \frac{r^2}{D} \left( \frac{\lambda}{2} + \frac{sr^2}{4} - \frac{r^4}{6} \right) \right), \quad (2)$$

where $N$ is the normalisation constant. Depending on the value of the two parameters, $\lambda$ and $D$, the distribution changes the shape from unimodal to bimodal behaviour, indicating a stochastic P-bifurcation which is caused by a noise-induced limit cycle of finite amplitude. The transition takes place if the exponent of the distribution $P$ develops an inflection point with regards to the variable $r$, i.e., at $(\ln P)' = 0$ and $(\ln P)'' = 0$

$$- \frac{D}{r} - \lambda r - sr^3 + r^5 = 0, \quad \frac{D}{r^3} - \lambda - 3sr^2 + 5r^4 = 0. \quad (3)$$

These two conditions constitute a set of polynomial equations in the variable $r^2$, which reduces to an expression for the independent variable

$$r^2 = \frac{9D + \lambda s}{6\lambda + 2s^2} > 0. \quad (4)$$

Eq.(4) shows that only values for $\lambda < 0$ are allowed, otherwise one would face unphysical imaginary radii. The resultant (in the mathematics literature called determinant) of the two polynomial equations (3) reads

$$\left( 3D + \lambda s + \frac{2s^3}{9} \right)^2 = \left( \lambda s + \frac{2s^3}{9} \right)^2 + \lambda^2 \frac{4(4\lambda + s^2)}{3}. \quad (5)$$

The bifurcation line in the $\lambda$-$D$ parameter plane which is explicitly determined by eq.(4) has the typical triangular shape with a base covering the deterministic bistability range and a cusp at $\lambda = - s^2/3, D = s^3/27$ (see figure 2).

FIG. 2: Stochastic bifurcation diagram of the stochastic differential equation without delay eq.(1), corresponding to eq.(5) for $s = 1$. The blue shaded area denotes the parameter values, where the probability distribution has a bimodal shape. The dashed line shows the border between the deterministic monostable (A) and the deterministic bistable (B) regime.

Hence, if we consider parameter values in the subthreshold regime, $-s^2/3 < \lambda < -s^2/4$, and gradually increase the strength of the noise, $D$, noise-induced oscillations at finite amplitude appear at an optimal noise strength, which is signalled by the second maximum in the stationary distribution. If the noise exceeds a critical value, these oscillations are finally wiped out again. Such features are quite common for noise-induced transitions in stochastic differential equations (see e.g. [18]).

It is the main purpose of the subsequent analysis to study how the classical scenario is affected by time-delayed feedback. We first focus on numerical simulations of eq.(1). We use a second order predictor-corrector scheme, along the lines of the Heun scheme (see [35] for related algorithms without time delay) with stepsize of about $10^{-3}$. Ensembles with about $10^9$ data points have been generated to evaluate the stationary behaviour. The data displayed in figure 3 show indeed that noise-induced bifurcations between unimodal and bimodal distributions occur in the system with delay when the noise strength is increased. Thus, the scenario known from the stochastic differential equation persists in the context of stochastic delay differential equations.

To illustrate the impact of time delay in a more systematic way let us focus on a parameter setup which gives rise to a bimodal distribution without delay and monitor how the shape of the distribution changes with $\tau$. Figure 4 clearly demonstrates that bimodality, i.e., the persistence of noise-induced oscillations is supported by integer time delay (in units of the intrinsic frequency $\omega$) (panel a) but that these oscillations are strongly suppressed at half integer time delay (panel b). In that respect the feedback in eq.(1) shares some similarity with time-delayed feed-
The linear deterministic part of eq.(1) induces a harmonic approximation is in excellent agreement with the simulation of an analytical approximation scheme, eq.(9). The analytical approximation is in excellent agreement with the simulation results.

FIG. 3: Stationary amplitude probability distribution of the stochastic delay differential equation for $\omega = 2\pi$, $\lambda = 0.18$, $s = 1$, $\tau = 0.25$, $K = 0.5$ and three values of the noise strength $D$ as obtained from numerical simulations of eq. (1) (solid lines). Dashed lines (white) show the respective result of an analytical approximation scheme, eq. (9). The analytical approximation is in excellent agreement with the simulation results.

We now consider parameter values close to such an instability. For that purpose we introduce a formal small parameter $\varepsilon$ which measures the distance from the instability. If scaled appropriately the time-dependent solution of the equation of motion (1) is then given by an amplitude modulated oscillation of the type $z(t) = \varepsilon A(\varepsilon^4 t) \exp(i\Omega_0 t)$. The appropriate scaling which balances the impact of all the different contributions in eq. (1) at an order $\varepsilon^5$ reads $\lambda - \lambda_0 = \delta\lambda = \varepsilon^4 \delta\lambda$, $s \to \varepsilon^2 s$, and $D \to \varepsilon^6 D$. If we use a simple Taylor series expansion for the delayed feedback term, $A(\varepsilon^4(t - \tau)) = A(\varepsilon^4 t) - \varepsilon^5 A'(\varepsilon^4 t) + \ldots$ then eq. (1) at leading order $\varepsilon^5$ results in

$$A'(\theta) = (\delta\lambda + s|A(\theta)|^2 - |A(\theta)|^4)A(\theta)$$

$$K \tau \exp(-i\Omega_0 \tau) A'(\theta) + \sqrt{2D} \exp(-i\Omega_0 \tau) \zeta(\theta),$$

where $\theta = \varepsilon^4 t$ denotes the slow time scale (and $'$ the corresponding derivative). If we recall that a phase transformation leaves an isotropic complex valued white noise invariant, we finally end up with

$$A'(\theta) = \frac{(\delta\lambda + s|A(\theta)|^2 - |A(\theta)|^4)A(\theta)}{1 + K \tau \exp(-i\Omega_0 \tau)} + \frac{\sqrt{2D} \zeta(\theta)}{|1 + K \tau \exp(-i\Omega_0 \tau)|}.$$  

Thus, close to an instability one is able to reduce the stochastic delay dynamics, eq. (1), to the ordinary stochastic differential equation (3) with Markovian property. Since eq.(3) has a phase symmetry, it is rather straightforward to write down the spherically symmetric stationary probability distribution (see, e.g., [34])

$$P(r) = N \tau \lambda \left( \frac{r^2}{D_{eff}} \left( \delta\lambda + s \frac{r^2}{4} - \frac{\theta^4}{6} \right) \right),$$

$$\delta\lambda = \lambda - K(1 - \cos(\Omega_0 \tau)),$$

$$D_{eff} = \frac{D}{1 + K \tau \cos(\Omega_0 \tau)}.$$  

Here we have again taken advantage of the spherical symmetry, which allows us to replace $A$ by $r \exp(i\phi)$.

Eq. (9) constitutes an analytical approximation for the stationary distribution of the stochastic time delay system, eq. (1). The expression differs from eq.(2), i.e. the stationary distribution of the system without time delay, by a shift in the parameter $\lambda$ and a rescaling of the noise, see eq.(9). The shift in the parameter $\lambda$ reflects the bifurcation structure of the deterministic time delay system (see appendix A). In addition, and somehow counterintuitively, the time delay has also an influence on the effective noise intensity in the system. Above all the impact of time delay can be made quantitative. The frequency $\Omega_0$ is determined by the implicit condition, eq. (9) as a function of $K$ and $\tau$. A priori the expression eq. (9) is valid close to an instability. However, the numerical results clearly demonstrate that the analytic expression is valid in a surprisingly large parameter region. Figure
clearly shows that numerical simulations and the analytical approximation are virtually indistinguishable for moderate values of the delay. Even for larger values of the delay, see figure 4, where deviations become noticeable the analytic expression still captures quite well the suppression of the bimodality of the stationary probability distribution at half integer delay times.

![Graph showing stationary amplitude probability distribution](image)

**FIG. 4**: Stationary amplitude probability distribution of the stochastic delay differential equation for $\omega = 2\pi$, $\lambda = -0.26$, $s = 1$, $D = 0.015$, $K = 0.5$, and different values of time delay: (a) integer time delay, (b) half integer time delay. Solid: results obtained from numerical simulations of eq. (1), dashed: corresponding analytical approximation, eq. (9).

Eq. (9) certainly fails if the effective noise strength changes sign and becomes negative. The corresponding mechanism, i.e., the sign of the expression $1 + K\tau \cos(\Omega_0 \tau)$, can in fact be related to the bifurcation diagram of the deterministic dynamical system, where the sign of this expression determines the character of the underlying Hopf bifurcation (see appendix A for the details). Interestingly, such a sub-supercritical transition was one of the key elements rectifying in misjudgements about time-delayed feedback control [30]. Overall, some care has to be applied when one uses the analytical expression, eq. (9), when $K\tau$ becomes larger than unity. In such cases it is easy to see that eq. (9) provides no longer a unique solution for $\Omega_0$. A “nontrivial” value enters the expression eq. (9) for which the effective noise strength remains to be positive. These “nontrivial” values for $\Omega_0$ are in fact directly related with the eigenvalue causing the Hopf instability on which the analytical approach has been based. As can be seen from figure 1 one again recovers quite a satisfactory analytical approximation even if the product $K\tau$ becomes larger than unity, and the approximation in fact seems to improve if one moves away from the “critical” case $K\tau = 1$. In summary, the stationary probability distribution of the stochastic time delay system and its analytical approximation reflects some of the fine structures of the underlying deterministic time delay dynamics.

Since the analytic expression, eq. (9), captures quite well the features of the stationary probability distribution, we may indeed base the computation of the stochastic bifurcation diagram on the analytic approximation. As eq. (9) differs from eq. (2) just by a delay dependent rescaling of the parameters the same is true for the stochastic bifurcation lines, see eq. (9). Figure 5 shows the result of such a straightforward analysis. Because of the scale invariance of eq. (5) it is in fact sufficient to constrain the analysis to the case $s = 1$. The most striking impact of the delay is a considerable shift of the bifurcation lines which is mainly periodic with regards to the time delay, together with a mild scaling of the triangular region.

## III. CORRELATION PROPERTIES AND COHERENCE RESONANCE

Coherence resonance is a dynamical phenomenon, which manifests itself in a nonmonotonic dependence of the correlation time upon the noise intensity, exhibiting a maximum at non-zero noise strength. There is also compelling evidence that coherence resonance is indicated by multimodality of stationary probability distributions and related stochastic bifurcations [18]. The model, eq. (1), is a suitable system to test such conjectures and to study in particular the impact of time delay on such scenarios. For that purpose we evaluate the autocorrelation function of the variable $x(t) = \text{Re}(z(t))$ and the corresponding power spectral density $I(\Omega)$. To begin we focus on numerical simulations of the stochastic time delay system eq. (1). We generate time traces of length $2^9\tau$ with step size $2^{-9}\tau$ resulting in a frequency resolution of about $\Delta\Omega \approx 10^{-2}$ with a cut-off frequency at about $\Omega_{\text{max}} \approx 10^3$. The statistical average is taken over an ensemble of 5000 time traces. In order to discount for linear resonance phenomena we monitor the spectrum normalised by the strength of the noise, $I(\Omega)/D$. In this way only effects caused by nonlinearities will show up in changes of the power spectrum.

To demonstrate coherence resonance and its relation with stochastic bifurcations we consider the dependence of the normalised spectrum on the noise strength for parameter variations, which do and which do not cross the stochastic bifurcation region, cf. figure 5. Figure 5(a) shows pronounced coherence resonance for noise amplitudes within the region of a bimodal stationary distribution. On increasing $D$ the additional power accumulates...
in the central peak, which increases by more than one order of magnitude compared to the background spectrum at an optimal noise amplitude \((D = 0.02)\). On increasing \(D\) further, the power then diffuses to the background spectrum and the central peak diminishes. The other parts of the spectrum are mainly unaffected by the change of the noise power. This coherence resonance phenomenon is not strictly linked with stochastic bifurcations since the same feature, even though less pronounced, is observed if we consider parameter variations which do not cross the region of bimodal stationary probability distributions, see figure 6(b).

As one can already guess from the structure of the bifurcation diagram of the deterministic model (see appendix A, figure 13) one expects a considerable impact of the time delay on the dynamics of the stochastic system. We are mainly interested in changes of the coherence resonance phenomenon. In order to focus on essential features we initially consider integer delays where shifts of the stochastic bifurcations are minimised, and on a noise strength close to the optimal value. As visible from the data displayed in figure 7 an increase of the time delay leads to a moderate sharpening of the central peak while on increasing the time delay further additional satellite lines occur. Such a feature is quite common in time-delayed feedback systems, and can be qualitatively attributed to a frequency filtering caused by the particular form of the time-delayed feedback term \([40]\).

The numerical findings presented so far just give a glimpse of all the phenomena that may appear in eq. 11. To gain deeper insight we need an analytic technique to approximate correlations in nonlinear stochastic time delay systems. For that purpose we adopt ideas that can be described as a variant of statistical linearisation and which have proven to be quite successful for analytical approximations of Fokker-Planck systems with nonlinear drift \([41]\). The main idea of the method, as applied to the nonlinear stochastic time delay system,
eq. (1), consists in approximating the nonlinear contribution \( |s(z(t))|^2 z(t) - |z(t)|^4 z(t) \) by an effective linear term \( \alpha z(t) \) resulting in a linear equation of motion

\[
\dot{z}(t) = (\tilde{\lambda} + i\omega)z(t) - K(z(t) - z(t - \tau)) + \sqrt{2D}\zeta(t),
\]

where

\[
\tilde{\lambda} = \lambda + \alpha \quad (11)
\]
denotes the rescaled bifurcation parameter. Contrary to a self-consistent mean field approach, which would replace \( |z(t)|^2 \) and \( |z(t)|^4 \) by its average values assuming some underlying Gaussian distribution (see e.g. [42]), statistical linearisation aims at an approximation which gives the best estimate of the nonlinear term by the linear one on average. Hence, the value of the coefficient \( \alpha \) is determined by minimising a suitable norm

\[
\langle |s(z(t)|^2 z(t) - |z(t)|^4 z(t) - \alpha z(t) |^2 \rangle \rightarrow \text{min.} \quad (12)
\]

In this way we obtain

\[
\alpha = \frac{s \langle |z(t)|^4 \rangle - \langle |z(t)|^6 \rangle}{\langle |z(t)|^2 \rangle}, \quad (13)
\]

where the averages are taken with regards to the stationary distribution.

The dynamics of the linear stochastic time delay system, eq. (10), which provides an analytic approximation to the nonlinear dynamics of eq. (1), can be solved in closed form (see e.g. the seminal contribution [43]). Linear equations can be conveniently solved in terms of eigenmodes. The characteristic equation corresponding to eq. (10) reads

\[
\Lambda = \tilde{\lambda} + i\omega - K(1 - \exp(-\Lambda\tau)) . \quad (14)
\]

One can solve this transcendental equation by using the different branches of the Lambert W-function [43]

\[
\Lambda_\ell = W_\ell[K\tau \exp(-(\tilde{\lambda} + i\omega - K)\tau)] + \tilde{\lambda} + i\omega - K, \quad \ell \in \mathbb{Z}. \quad (15)
\]

The time-dependent solution of eq. (10) is then expressed as a superposition of eigenmodes

\[
z(t) = \sum_\ell C_\ell(t), \quad (16)
\]

where the time evolution of the coefficients is governed by a variation of constants formula

\[
\dot{C}_\ell(t) = \Lambda_\ell C_\ell(t) + \frac{\sqrt{2D}\zeta(t)}{N_\ell}, \quad (17)
\]

with the weight

\[
N_\ell = 1 + K\tau \exp(-\Lambda_\ell \tau) \quad (18)
\]

following from the normalisation of eigenmodes. Solving eq. (17) the stationary solution follows from eq. (16)

\[
z(t) = \sqrt{2D} \int_0^\infty T(t')\zeta(t - t')\,dt', \quad (19)
\]

where we have introduced the propagator of the linear equation (10)

\[
T(t) = \sum_\ell \frac{\exp(\Lambda_\ell t)}{N_\ell}, \quad (t > 0) . \quad (20)
\]

Using the correlation properties of the normalised complex valued noise, \( \langle \zeta(t)\zeta(t') \rangle = 0 \) and \( \langle \zeta(t)\zeta(t') \rangle = 2\delta(t - t') \), we easily conclude that the autocorrelation functions of the real part \( x(t) = (z(t) + z^*(t))/2 \) and imaginary part \( y(t) = (z(t) - z^*(t))/(2i) \) can be written as

\[
\langle x(t)x(0) \rangle = \langle y(t)y(0) \rangle = \frac{1}{2} \text{Re} \langle z(t)z^*(0) \rangle . \quad (21)
\]

Using again the correlation properties of the noise the remaining autocorrelation function follows from eq. (19) and eq. (20) after some short algebra

\[
\langle z(t)z^*(0) \rangle = 4D \sum_{\ell,\ell'} \frac{\exp(\Lambda_{\ell\ell'})}{N_\ell N_{\ell'}(-\Lambda_{\ell} - \Lambda_{\ell'})}. \quad (22)
\]

For the power spectral density, i.e., for the Fourier transform of eq. (21) we finally arrive at

\[
I(\Omega) = 2D\text{Re} \sum_{\ell,\ell'} \frac{-2\Lambda_{\ell\ell'}}{(\Omega^2 + \Omega^2)N_\ell N_{\ell'}(-\Lambda_{\ell} - \Lambda_{\ell'})}. \quad (23)
\]

As expected, the power spectral density essentially consists of Lorentzian lines where the linewidth is determined by the real part of the eigenvalues, eq. (15). One can express the series, eq. (23), in closed analytic form. For that purpose we recall that the propagator, eq. (20), obeys the deterministic equation of motion, eq. (10), with \( D = 0 \), with initial condition \( T(0) = 1 \) and \( T(\theta) = 0 \), \( -\tau < \theta < 0 \). Hence, the Laplace transform of this solution reads

\[
\tilde{T}(\Omega) = \sum_{\ell} \frac{1}{i\Omega - (\Lambda_\ell + i\omega - K) - K \exp(-i\Omega)} . \quad (24)
\]

It is now rather straightforward to show that the sum in eq. (24) can be essentially expressed as the absolute square of the Laplace transform, eq. (24), so that we arrive at the closed analytic formula [43]

\[
I(\Omega) = D(|\tilde{T}(\Omega)|^2 + |\tilde{T}(-\Omega)|^2). \quad (25)
\]
The data shown in figure 7 prove the accuracy of the analytical approximation, which is able to reproduce the numerical results within 1% of accuracy over a rather wide generic range of parameter values. Within our level of resolution the analytic and the numerical values can hardly be distinguished and one can safely use the analytic formula to estimate coherence resonance signatures in our model. To some extent the success of the approximation depends crucially on the correct estimate of the effective coefficient by eq. (13).

To demonstrate the modulation of coherence resonance caused by the time delay, we use the correlation time, which in a linear approximation is proportional to the inverse of the width $\Gamma$ of the main spectral peak. Alternatively, the correlation time may be introduced as in [46] by integrating over the absolute value of a normalised autocorrelation function, i.e.

$$ t_{\text{cor}} = \frac{1}{\langle x(0)x(0) \rangle} \int_0^\infty |\langle x(t)x(0) \rangle|dt. \quad (26) $$

The correlation function can be evaluated using eqs. (21) and eq. (22). To deal with the modulus we approximate the fast oscillating part of the autocorrelation function by $\pi^{-1} \int_{-\pi/2}^{\pi/2} \cos(\phi)d\phi = 2/\pi$ [40]. Thus we arrive at

$$ t_{\text{cor}} = \frac{2}{\pi \langle x(0)x(0) \rangle} \sum_{\ell \ell'} \left| N_\ell N^*_\ell' (\Lambda_\ell - \Lambda_{\ell'}) \right| \text{Re}(\Lambda_\ell). \quad (27) $$

For the numerical evaluation of eq. (26) we simulated an ensemble of 300 time traces of length $10^4$ with about $5 \times 10^5$ data points, to compute the autocorrelation function and the integral of its absolute value. The simulations were repeated for 100 different values of the noise intensity in the interval $D \in [0.001, 0.1]$. The comparison between the simulation and the analytical estimate is shown in figure 8. Note that for the analytical results we only take the main branch of the Lambert W-function ($\ell = \ell' = 0$) in eq. (27) into account. For small time delays the main branch dominates the dynamics and adding more branches does not cause significant changes of the result. Our approximation is in good agreement with the numerics for integer delay times.

The enhancement and the suppression of the correlation time is caused by the stability of the deterministic focus: it shows non-monotonic behaviour as a function of $\tau$, cf. figure 13 (App.A) for the bifurcation diagram of the underlying deterministic system. The focus becomes more stable for half-integer delays and less stable for integer ones. Furthermore, an increase of the integer delays shifts the optimal noise intensity to higher values and the correlation becomes stronger. The shift of the optimal noise intensity can be explained by the fact that the regime of bimodal probability distribution is changed by the delay time, which is shown in figure 5(b). For a fixed value of $\lambda$ a higher value of the noise is necessary to reach the regime of bimodality, see figure 5(b).

The enhancement and the suppression of the correlation time is clearly visible in figure 8. The maximum value of the correlation time at a certain optimal noise intensity signals coherence resonance. As one would expect from the results in section II, coherence resonance is suppressed for half-integer delay times and enhanced for integer ones. For a fixed value of $\lambda$ a higher value of the noise is necessary to reach the regime of bimodality, see figure 5(b).
multimodality have been considered as a key characteristic of coherence resonance, but such a viewpoint is not entirely satisfactory. In [18] it is pointed out that the highest degree of regularity at a certain optimal noise intensity, i.e., pronounced coherence resonance, is provided by a stationary distribution with bimodal shape. But outside of the bimodal regime less pronounced coherence resonance can still be observed. Hence, it is tempting to figure out features within the stationary distribution that can be used to detect coherence resonance. Within our model the mechanism of coherence resonance is related with the ghost of the stable periodic orbit that is created at \( \lambda = -s^2/4 \) and \( |z| = \sqrt{s/2} \) in the deterministic system without time delay. To measure the significance of such a phase space structure we suggest to consider the cumulative probability of states having amplitudes beyond the critical radius \( |z| = \sqrt{s/2} \) at which the pair of stable and unstable periodic orbits is created
\[
    g(D) = \int_{-\infty}^{\infty} P(r) dr .
\]
Consequently, the quantity \( g(D) \) is the weight of the probability distribution above the radius of the ghost limit cycle induced by noise. Therefore, we call it ghost weight. The ghost weight can even be evaluated by analytical means using eq. (19). It is essentially a monotonically increasing function as the noise intensity is spreading out the probability in phase space. The resonance characteristics, i.e., the dynamical impact of the phase space structure close to the saddle-node bifurcation point, is indicated by the increment of the function, i.e., by the sensitivity of \( g(D) \) with respect to \( D \). Indeed, the derivative \( dg/dD \) is able to show quite accurately the resonance, see figure [9].

That property seems to be quite generic, as it holds in a rather large part of the parameter region, see figure [10]. While the numerical values of the correlation time \( t_{cor} \) and of \( dg/dD \) certainly differ (they do not even have the same units) the resonance characteristics, i.e., the maximum visible in the ridge of the two functions coincide quite nicely. The derivative of the ghost weight together with the correlation time can be used to detect and hence, provides an explanation of the mechanism of coherence resonance outside the regime of bimodality.

In summary, the techniques behind eq. (25) provide powerful tools to estimate correlation properties of nonlinear stochastic time delay systems.

IV. CONCLUSION

We have presented an extensive study of coherence resonance in non-excitatory systems with a particular focus on the impact of time-delayed feedback. For that purpose we introduced a paradigmatic model system, i.e., a Hopf normal form subjected to noise and a time-delayed feedback coupling. That setup allowed us to study in detail the coherence resonance phenomenon related to the saddle-node bifurcation of periodic orbits which is connected with a subcritical Hopf bifurcation. Given that normal forms are able to capture features of general dynamical systems we propose that our findings have a
We have applied analytical tools to investigate stochastic time delay equations and to confirm the numerical findings of our model. For the stationary probability distribution a centre manifold analysis has been developed to predict and confirm the shape of the stationary probability distribution with a particular focus on detecting stochastic bifurcations. While a priori such analytic approximations are only valid close to bifurcation points we find that our expression, eq.(9), is able to capture even quantitatively the behaviour of the stationary state in a large region of the parameter space. Time delay leads to a characteristic modulation and scaling of the stochastic bifurcation pattern as visible in figure 5. That pattern can be captured by parameters renormalised by the time delay, see eq.(9).

To investigate correlation properties we have used the fact that linear stochastic time delay systems, like their deterministic counterpart, admit a complete analytic solution. Adapting the concept of stochastic linearisation, which is known to deliver rather accurate analytic approximations for Fokker-Planck equations with nonlinear drift, we have derived closed expressions for the autocorrelation function as well as the power spectral density. Such an approximation scheme captures the numerical findings quite well, in fact, within 1% of the numerical values. There is excellent agreement between the numerical results obtained from the power spectral density and the analytical estimates (see, e.g., figure 7). Thus we are able to evaluate correlation properties in detail.

The maximum of a correlation time is often used as an indicator for coherence resonance. The correlation time may be defined in two different ways: either a direct calculation from the autocorrelation function, or the inverse of the linewidth in the power spectral density. The latter is analytically accessible from the eigenvalues, eq.(10). The correlation time shows a maximum, i.e., the linewidth shows a minimum, which indicates coherence resonance even if the parameter setting does not correspond to a bimodal stationary probability distribution, see figure 9. Therefore the ghost weight, which is calculated from the probability distribution, but does not depend on its shape, has been introduced as a signature of coherence resonance. In particular, it allows to draw an analogy between the mechanism of coherence resonance in excitable systems based on two competing time-scales and in non-excitable systems, where this phenomenon occurs due to the competition between the noisy focus and the ghost of the limit cycle. Hence, coherence resonance and noise-induced bifurcation are two slightly different facets of stochastic time delay systems.

The investigation of our simple model systems and the tools and signatures that have been developed in this context may help to broaden our understanding of coherence resonance in stochastic delay systems, and to open up insight into the dynamical interplay between stochastic inputs and time delayed interactions.

Acknowledgments

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Appendix A: Bifurcation analysis of the deterministic model

The bifurcation analysis of the deterministic model, i.e., eq.(1) with $D = 0$, can be performed to a large extent by analytical means. The linear stability of the trivial fixed point $z = 0$ is obviously governed by the characteristic equation

$$\Lambda = \lambda + i \omega - K(1 - \exp(-\Lambda \tau)). \quad (A1)$$

If the eigenvalue becomes purely imaginary, $\Lambda = i \Omega$, a Hopf bifurcation takes place. Thus, eq.(A1) yields (see also eq.(9))

$$\Omega = \omega - K \sin(\Omega \tau), \quad (A2)$$

$$\lambda = K(1 - \cos(\Omega \tau)) \quad . \quad (A3)$$

We discuss the bifurcation diagram in a $(\lambda, K)$ parameter plane, for fixed value of the delay $\tau$. We first uncover the structure of the various bifurcation lines in the $(\lambda, K)$ parameter plane and translate the results to the $(\lambda, \tau)$ plane at the end of this appendix (see figure 13).

Eq.(A1) yields for the bifurcation lines the parametric representation

$$K = \frac{\omega - \Omega}{\sin(\Omega \tau)}, \quad (A4)$$

$$\lambda = K(1 - \cos(\Omega \tau)), \quad \Omega \tau \in (n \pi, (n+1) \pi).$$

Generically, eq.(A4) results in a single monotonic branch and a nested structure of parabolic branches (see figure 11). As for the direction of the bifurcation one may compute the change of eigenvalue with respect to $\lambda$. Implicit differentiation of eq.(A4) results in

$$\operatorname{Re} \left( \frac{d\Lambda}{d\lambda} \right) \bigg|_{\lambda = i \Omega} = \frac{1 + K \tau \cos(\Omega \tau)}{|1 + K \tau \exp(i \Omega \tau)|^2}. \quad (A5)$$

Since the extrema of the bifurcation lines, eq.(A4), are determined by the condition $1 + K \tau \cos(\Omega \tau) = 0$ we conclude that the real parts become positive when crossing the left part of one of the parabolic branches and negative when crossing the right branch (see figure 11). Finally, to evaluate the type of Hopf bifurcation one would need to evaluate the cubic coefficient of the corresponding normal form. It is rather straightforward but slightly tedious to show that the real part of the cubic coefficient is given by $s(1 + K \tau \cos(\Omega \tau))$ (see appendix B). Hence, the Hopf bifurcations on the single monotonic branch as well as on the left part of a parabolic branch are subcritical while those on the right part are supercritical.
saddle-node bifurcation. Thus, to label this type of bifurcation we call it a rigid high symmetry of our equations of motion, see eq.\((A7)\). The second mechanism for the creation of a pair of harmonic solutions, i.e., for a saddle-node bifurcation, is caused by eq.\((A7)\). On changing the parameter \(K\) a real pair of \(\alpha\) values can be created resulting in two harmonic solutions of different period. The corresponding condition for this type of bifurcation follows from the derivative of eq.\((A7)\) and obviously reads
\[
1 + K\tau \cos(\alpha\tau) = 0.
\] (A9)

Eqs.\((A6)\) and \((A9)\) determine horizontal straight bifurcation lines which precisely hit the the extrema of the Hopf and the rigid saddle-node bifurcation lines (see figures \[11\], \[12\]). Since we still have to ensure that eq.\((A6)\) gives real solutions for the amplitude the line does not extend to infinity. In fact, the bifurcation line determined by eq.\((A9)\) starts at the sub-supercritical transition with a small value for \(|A|\). It extends to the left until \(\lambda = -s^2/4\) where it hits the extremum of the corresponding rigid saddle-node bifurcation line, and where eq.\((A9)\) provides a doubly degenerate solution. Hence, on increasing \(\lambda\) we obtain another branch of the saddle-node bifurcation line with large amplitudes \(|A| \geq s/2\) which then extends from \(\lambda = -s^2/4\) to infinity (see figure \[12\]). As this saddle-node bifurcation results in a pair of periodic orbits which attain different periods we label the saddle-node bifurcation as non-rigid.

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In a similar way (cf. eqs. (A10) and (A11)) we derive a closed analytic formula for the rigid saddle-node bifurcation line

$$\tau_{sn} = \pm \arccos \left( \frac{(K - \lambda - s^2/4)/K + 2\pi n}{\omega \pm \sqrt{(4K - 4\lambda - s^2)(4\lambda + s^2)/4}} \right), \quad n \in \mathbb{N}. \tag{A11}$$

Figure 13 shows the result of this analysis for the parameter setting used in sections II and III, and thus allows a clear understanding of the impact of time delay.

![Figure 13: Hopf curves and saddle-node bifurcation curves in the $(\lambda, \tau)$ plane, corresponding to eq. (A10) and eq. (A11), for $\omega = 2\pi$, $K = 0.5$, $s = 1$, and $n = 0, \ldots, 6$. Blue (black): Hopf bifurcation curves $\tau_{h}$ (solid: supercritical bifurcation, dashed: subcritical bifurcation), bronze (grey): saddle-node bifurcation of limit cycles curves $\tau_{sn}$ (solid: rigid bifurcation, dashed: non-rigid bifurcation). The light blue shaded area indicates the region with stable trivial fixed point and without any harmonic periodic orbit. It is of interest for coherence resonance (see figure 1).](image)

**Appendix B: Centre manifold reduction of the stochastic delay differential equation**

To perform a centre manifold reduction we consider a parameter value $\lambda = \lambda_0 + \delta \lambda$ close to a Hopf bifurcation, i.e., $\lambda_0$ obeys eq. (A1) with eigenvalue $\Lambda = i\Omega_0$ (cf. eq. (6) as well). The formal calculation follows standard procedures [36]. We here follow the approach proposed in [41]. Using the condition on the parameter $\lambda$ we rewrite eq. (A1) as

$$\dot{z}(t) = (\lambda_0 - K)z(t) + Kz(t - \tau) + f(t), \tag{B1}$$

where the inhomogeneous part, given by

$$f(t) = (\delta \lambda + s|z(t)|^2 - |z(t)|^4)z(t) + \sqrt{2D_\zeta(t)} \tag{B2}$$

contains all the contributions of the unfolding which are supposed to be all of the same small order. To perform a centre manifold reduction recall that the phase space of a delay differential equation is given by a function space, here on $[-\tau, 0]$. Such a feature is captured by introducing the notation $Z_i(\theta) = z(t + \theta)$ with the delay variable $\theta \in [-\tau, 0]$. Since the solution of the linear part of eq. (B1) reads $Z_i(\theta) = \exp(i\Omega_0 t)\exp(i\Omega_0 \theta)$ the centre manifold is tangential to $V(\theta) = \exp(i\Omega_0 \theta)$ and can be written as

$$z(t + \theta) = C(t)\exp(i\Omega_0 t) + R(t, \theta). \tag{B3}$$

Here, $C(t)$ denotes the complex coordinate on the centre manifold and $R(t)$ captures the higher order terms of an expansion in $C(t)$. To define the coordinate on the centre manifold uniquely one needs to impose a constraint on the higher order terms. It is quite convenient to work with an orthogonality condition with regard to the adjoint eigenmode of the linear system [41]. Such a constraint results in the condition

$$0 = R(t, 0) + K \int_{-\tau}^{0} \exp(-i\Omega_0(\theta + \tau))R(t, \theta)d\theta. \tag{B4}$$

Combining eqs. (B3) and (B4) we obtain for the coordinate

$$NC(t) = z(t) + K \int_{1-\tau}^{t} \exp(i\Omega_0(t - \theta - \tau))z(\theta)d\theta \tag{B5}$$

with the normalisation $N = 1 + K\tau\exp(-i\Omega_0\tau)$. Taking the time derivative and using eq. (B1) we arrive at the equation of motion on the centre manifold

$$\dot{C}(t) = i\Omega_0C(t) + f(t)/N. \tag{B6}$$

If we evaluate the expression (B2) on the centre manifold (B3) to lowest order we thus end up with (cf. eq. (7))

$$\dot{C}(t) = i\Omega_0C(t) + \frac{(\delta \lambda + s|C(t)|^2 - |C(t)|^4)C(t)}{N} + \frac{\sqrt{2D_\zeta(t)}}{N}. \tag{B7}$$

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