Prenatal exposure to multiple metallic and metalloid trace elements and the risk of bacterial sepsis in extremely low gestational age newborns: A prospective cohort study

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Abstract

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Author contributions

CB conceptualized the study, conducted all statistical analyses, created all tables and figures, and wrote the original draft of the manuscript. LS and AG prepared umbilical cord biospecimens for metals assessment. AG and PP developed the ICP-MS/MS methodology, measured metal concentrations, and provided critical review of the manuscript. KR curated the data. LE provided critical review of the manuscript. TO’S and RF acquired the funding, provided project administration, and provided critical review of the manuscript. All authors contributed to the article and approved the submitted version.

Conflict of interest

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

Ethics statement

The ELGAN study was approved by the Institutional Review Boards at each of the 14 participating study sites and informed consent was provided by all enrolled mothers for themselves and their infants.

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Supplementary material

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fepid.2022.958389/full#supplementary-material
**Background:** Prenatal exposures to metallic and metalloid trace elements have been linked to altered immune function in animal studies, but few epidemiologic studies have investigated immunological effects in humans. We evaluated the risk of bacterial sepsis (an extreme immune response to bacterial infection) in relation to prenatal metal/metalloid exposures, individually and jointly, within a US-based cohort of infants born extremely preterm.

**Methods:** We analyzed data from 269 participants in the US-based ELGAN cohort, which enrolled infants delivered at <28 weeks’ gestation (2002–2004). Concentrations of 8 trace elements—including 4 non-essential and 4 essential—were measured using inductively coupled plasma tandem mass spectrometry in umbilical cord tissue, reflecting in utero fetal exposures. The infants were followed from birth to postnatal day 28 with bacterial blood culture results reported weekly to detect sepsis. Discrete-time hazard and quantile g-computation models were fit to estimate associations for individual trace elements and their mixtures with sepsis incidence.

**Results:** Approximately 30% of the extremely preterm infants developed sepsis during the follow-up period (median follow-up: 2 weeks). After adjustment for potential confounders, no trace element was individually associated with sepsis risk. However, there was some evidence of a non-monotonic relationship for cadmium, with hazard ratios (HRs) for the second, third, and fourth (highest) quartiles being 1.13 (95% CI: 0.51–2.54), 1.94 (95% CI: 0.87–4.32), and 1.88 (95% CI: 0.90–3.93), respectively. The HRs for a quartile increase in concentrations of all 8 elements, all 4 non-essential elements, and all 4 essential elements were 0.92 (95% CI: 0.68–1.25), 1.19 (95% CI: 0.92–1.55), and 0.77 (95% CI: 0.57–1.06). Cadmium had the greatest positive contribution whereas arsenic, copper, and selenium had the greatest negative contributions to the mixture associations.

**Conclusions:** We found some evidence that greater prenatal exposure to cadmium was associated with an increased the risk of bacterial sepsis in extremely preterm infants. However, this risk was counteracted by a combination of arsenic, copper, and selenium. Future studies are needed to confirm these findings and to evaluate the potential for nutritional interventions to prevent sepsis in high-risk infants.

**Keywords**

metals; mixtures; pregnancy; sepsis; preterm (birth)

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**Background**

Sepsis, an extreme immune response to an infection of the bloodstream, is a leading cause of neonatal morbidity and mortality. In the last decade, there were an estimated 3,930 cases of neonatal sepsis per 100,000 live births worldwide and nearly 20% of cases were fatal (1). While proximal risk factors such as preterm labor, prolonged rupture of membranes, and intra-amniotic infections have been recognized (2), preventing neonatal sepsis remains a key challenge (3). Identifying more distal risk factors for neonatal sepsis that are modifiable may be a promising strategy, as it is often too late to prevent fetal and neonatal infections by the time medical attention is sought (3). One such plausible risk factor is prenatal exposure to environmental metallic and metalloid trace elements, which are ubiquitous in the air, soil, water, and food supply (4). Several non-essential trace elements, including arsenic (a
metalloid) (5), cadmium (a metal) (6), lead (a metal) (7), and mercury (a metal) (8), are recognized as immunotoxicants. In contrast, other nutritionally-essential trace elements such as copper (a metal) (9), manganese (a metal) (10), selenium (a metalloid) (11), and zinc (a metal) (10) may be necessary for optimal immune function, with experimental studies in animals indicating adverse effects from deficiencies and excesses (12–15).

Non-essential metallic and metalloid trace elements have distinct mechanisms of immunotoxicity, with many implicated as immunosuppressants, particularly when exposures occur early in life. For example, arsenic exposures in childhood have been linked to reductions in the proportion of CD4 (“helper”) T cells in peripheral blood (16), which capture and kill bacteria from infected dendritic cells. Prenatal exposures to cadmium have also been tied to alterations in lymphocyte composition in infancy (17), and prenatal methylmercury exposures to alterations that persist into childhood (18). At the same time, non-essential metals and metalloids could promote inflammation, which is an appropriate response to the presence of infection but can be dangerous when excessive (19). Research on lead has identified links between in utero exposures with lower concentrations of the anti-inflammatory cytokines interleukin (IL)-4 and IL-8, as well as higher concentrations of the pro-inflammatory cytokine tumor necrosis factor alpha (TNF-α) in early childhood (20), suggesting exposures may have implications for infection elimination and/or tissue damage (19). In the United States, over 85% of women of reproductive age have measurable levels of arsenic, cadmium, lead, and mercury in their bodies (21), making immunotoxic effects of prenatal exposures a pertinent concern.

It has been posited that non-essential metals and metalloids gain entry into human cells by “mimicking” and potentially competing with nutritionally-essential trace elements (22). For this reason, epidemiologic studies of exposure mixtures encompassing both non- and nutritionally-essential trace elements have become popular under the premise that nutritional interventions could counteract environmental metal/metalloid toxicity (23). Although few epidemiologic studies to date have considered the role of complex trace element mixtures in the development of immune-mediated and infectious diseases (23), nutritionally-essential elements are recognized for their important roles in the immune system. Zinc, for example, modulates the inflammatory response by binding to and inhibiting IKK-β, a protein important for cytokine-activated intracellular signaling. In a mouse model of polymicrobial sepsis, a zinc-deficient diet promoted insufficient control of IKK-β, leading to excessive inflammation (24). Therefore, we hypothesized that greater prenatal exposures to nutritionally-essential trace elements, namely copper, manganese, selenium, and zinc would be associated with a reduced risk of neonatal sepsis, whereas the opposite would be true for the non-essential trace elements arsenic, cadmium, mercury, and lead (i.e., greater prenatal exposures to non-essential trace elements would be associated with an increased risk of neonatal sepsis). To test these hypotheses, we analyzed each element individually and complemented this approach with quantile g-computation to assess how combinations of exposures in utero influences the risk of sepsis within the first few weeks of life. We conducted all analyses within the Extremely Low Gestational Age Newborns (ELGAN) cohort, which is comprised of infants born extremely preterm who are at a markedly increased risk of developing neonatal bacterial sepsis due to having immature immune systems and requiring invasive medical procedures and devices (25).
Materials and methods

The ELGAN cohort

The ELGAN study is an ongoing, multi-center, prospective cohort originally designed to assess contributors to brain damage in extremely preterm infants (born prior to 28 weeks’ gestation) (26). Infants were enrolled from 2002 to 2004 from 14 study sites across five states: Connecticut, Illinois, Massachusetts, Michigan, and North Carolina. The institutional review boards of each of the participating institutions approved enrollment and consent procedures. Of the 1,506 infants initially enrolled, umbilical cord tissue samples used to quantify prenatal metal and metalloid exposures were available for a subset of \( n = 292 \). However, 12 of the specimens had insufficient sample mass and another 11 infants were excluded for missing data on relevant covariates (detailed further below), bringing the final analytic sample size to \( n = 269 \).

Assessment of prenatal metal and metalloid trace element exposures in umbilical cord tissue

Study pathologists collected umbilical cord tissue specimens according to a standardized protocol. Briefly, using a sterile technique and scissors, two 1-cm segments of tissue were collected from the extra-abdominal umbilical cord attached to each infant within 1-h of being clamped and cut following delivery. The segments were immediately placed into separate cryovials, immersed into liquid nitrogen, and transported to a −80°C freezer for long-term storage. The samples were later shipped on dry ice to the Wadsworth Center for quantification of metal and metalloid trace elements.

Upon receipt at the Wadsworth Center, the umbilical cord samples were checked against the shipping manifest, accessioned per standard laboratory procedure, and stored at −80°C pending analysis. Samples were later thawed and, if applicable, sectioned into segments for analysis with a final dry mass of ~50 mg. Samples were sectioned using high-purity tantalum tools that were fabricated in-house to avoid contamination with other trace elements typically found in stainless steel. All samples were then rinsed with double-deionized (DDI) water to remove superficial blood and placed in a 13-ml acid-washed tube before being freeze-dried to constant mass using a slow 5-step program to ensure thorough removal of water content. The samples were digested in concentrated double-distilled HNO\(_3\) using a Microwave Assisted Reaction System (MARS) 6 equipped with Xpress vessels via the “One-Touch Animal Tissue” method. Digests were diluted to 10-mL with DDI water and stored at 4°C pending analysis. Sample digests were analyzed using an Agilent 8900 Inductively Coupled Plasma tandem Mass Spectrometer (ICP-MS/MS) equipped with a SPS 4 autosampler and an Octopole Reaction System (ORS) with axial acceleration technology. The method was optimized for analyzing umbilical cord tissue samples while reducing polyatomic and isobaric interferences. The method was validated against four Standard Reference Materials (SRM) obtained from the National Institute of Standards and Technology (NIST), Gaitherburg, MD: NBS 1577 Bovine Liver, NIST 1577b Bovine Liver, NIST 1577c Bovine Liver, and NIST 8414 Bovine Muscle Powder. The SRMs were freeze-dried, digested, and analyzed in the same manner as the cord samples. Values for sample spikes, duplicates, blanks, and calibrators were carefully monitored throughout the
study. A subset of sample digests were analyzed in duplicate \((n = 75)\) or triplicate \((n = 23)\) consistent with the Wadsworth Center’s quality management system; the final trace element concentration was reported as the average of the replicate values. The method limits of detection (LOD) for the 8 trace elements were: arsenic \((0.42 \text{ ng/g})\), cadmium \((0.32 \text{ ng/g})\), copper \((0.074 \mu\text{g/g})\), lead \((2.6 \text{ ng/g})\), manganese \((0.010 \mu\text{g/g})\), mercury \((0.79 \text{ ng/g})\), selenium \((0.10 \mu\text{g/g})\), and zinc \((1.4 \mu\text{g/g})\). All umbilical cord tissue samples had detectable levels of arsenic, copper, lead, manganese, selenium, and zinc. For mercury and cadmium, one and five samples respectively were at or below the detection limit so concentration values were replaced by the corresponding limit divided by \(\sqrt{2}\).

**Diagnostic testing for bacterial sepsis**

After delivery, trained research nurses used structured data collection forms to review medical charts. As previously described by Patel et al. (27) infant bacterial blood cultures were reported on the forms for days 7, 14, 21, and 28. As a result, the specific day on which the blood harbored an organism is unknown, only the week during which bacteremia first occurred. Incident bacterial sepsis was defined as the first instance of bacteremia.

Neonatal sepsis is often characterized by the timing of onset with cases presenting at <7 days of life referred to as early-onset and afterwards as late-onset (25). This characterization is based on suspected etiology. Early-onset cases are generally considered to be the result of pathogen transmission from the female genitourinary tract, infecting the fetus *in utero* or during delivery. In contrast, late-onset cases are hospital- or community-acquired. The bacteria responsible for the initial infection tends to differ by timing of onset as well. The majority of early-onset cases are attributable to *Escherichia coli* and group B streptococcus account whereas late-onset cases are mostly driven by Coagulase-Negative Staphylococci and *Staphylococcus aureus* (28, 29). The ELGAN study did not collect information on the bacterial species isolated from cultures (27). However, we divided sepsis cases as early- or late-onset to examine potential differences in metal/metalloid-associated risk profiles.

**Covariates**

We considered adjusting for many factors that might confound associations of prenatal metal and metalloid trace element exposures with neonatal bacterial sepsis. We narrowed these down by using a directed acyclic graph (DAG) to identify variables associated with both metal/metalloid exposures and sepsis, excluding any that could be causal intermediates (Supplementary Figure 1) (30). In all models, we adjusted for the geographic region in which the infant was delivered as well as maternal race/ethnicity, age, educational attainment, Medicaid coverage (a proxy for income), pre-pregnancy body mass index (BMI), and smoking status during pregnancy. Maternal sociodemographic characteristics pre-pregnancy BMI were self-reported shortly after delivery. Pregnancy and infant characteristics were abstracted from medical records.

**Statistical analyses**

All statistical analyses were conducted in R (version 4.0.3) (31).
For each trace element, we first calculated measures of central tendency and distributions, both overall and stratified by covariates. Most of the trace elements had right-skewed distributions, particularly the non-essential elements. Therefore, we calculated Spearman’s rank coefficients ($\rho$) to evaluate how the metals and metalloids were correlated with one another.

To conduct a time-to-event analysis of incident bacterial sepsis, which was ascertained on a weekly basis, the data were transformed into a person-period structure (32). In its original form, the data were organized such that each infant was represented by one row; after transformation, each infant had multiple rows reflecting each week they were under observation. Infants who were transferred to another hospital (and therefore lost to follow-up) or who did not develop sepsis during the first 28 days of life were right-censored.

To analyze the trace elements individually, we fit separate discrete time hazard models to estimate hazard ratios. The models were based on the probability of developing bacterial sepsis each week, conditional on being sepsis-free the week prior. Conditional probabilities were estimated using a complementary log-log link function with four baseline hazards (i.e., one for each week of follow-up time) (33). Indicator variables were used to model quartiles of cord metal/metalloid concentrations (relative to the lowest quartile) after adjustment for potential confounders. Geographic region (the Midwest, New England, or North Carolina), race/ethnicity (Hispanic, non-Hispanic black, non-Hispanic white, non-Hispanic other), educational attainment level (less than a high school diploma, high school diploma, or college degree), Medicaid coverage, and smoking status during pregnancy (yes or no) were modeled using indicator variables whereas age at delivery (years) and pre-pregnancy BMI ($kg/m^2$) were modeled continuously. Because recruitment into the study was based on extremely low gestational ages and, consequently, many of the infants were multiples (e.g., twins or triplets), we computed cluster-robust standard errors via bootstrapping at the pregnancy level with 1,000 iterations (34). We tested for deviations from the proportional hazards assumption by using likelihood ratio tests to compare models with and without interaction terms for metals/metalloids and time. To test for linear relationships between cord concentrations and sepsis incidence, we modeled the median trace element concentration in each quartile as a continuous variable (35); we considered the corresponding p-value as indicative of a linear trend if it was below 0.05.

To analyze trace element mixtures, we used quantile g-computation (36). These models were fit similarly to the individual element models in that we again used a complementary log-link function with four baseline hazards and covariates to estimate adjusted discrete-time hazard ratios. However, instead of estimating hazard ratios for each metal/metalloid individually, quantile g-computation estimates a hazard ratio for the exposure mixture. This hazard ratio can be interpreted as the change in the log-hazard of sepsis per a simultaneous one-quartile increase of all elements in the mixture. We tested three sets of mixtures: (1) all 8 trace elements combined; (2) the 4 non-essential trace elements (arsenic, cadmium, lead, mercury); and (3) the 4 essential trace elements (copper, manganese, selenium, zinc). For the latter two sets, the trace elements not included in the specific exposure mixture of interest were still included as model covariates so as to control for potential confounding by co-exposures. Adjustments were also made for the same model covariates included in the
individual trace element models. In each of the quantile g-computation models, the mixture components were assigned a weight that can be either negative or positive in direction. If all assigned weights are in the same direction, they are constrained to sum to one and can be interpreted as the proportional contribution of a given trace element to the overall association with incident sepsis. If the assigned weights have different directions, their absolute values are instead constrained to sum to two, constituted by one for the sum of all positive weights and one for the sum of all negative weights, and can be interpreted as the proportional contribution of a given trace element to the positive (or negative) partial association with incident sepsis. Cluster-based bootstrapping with 1,000 iterations was performed to estimate standard errors that appropriately account for the multiple pregnancies in the dataset.

Sensitivity and stratified analyses

For both individual and trace element mixtures, we attempted to re-fit the models stratified by early- or late-onset, but the number of early-onset cases (n = 19) was too small for reliable estimation. Instead, we performed analyses restricted to late-onset cases to explore whether there were any differences in metal/metalloid associations from the models of sepsis overall.

Because the ELGAN cohort sampled participants born extremely preterm at <28 weeks’ gestation, our analyses were implicitly partially adjusted for gestational age. However, as depicted in our DAG (Supplementary Figure 1), gestational age was considered to be a consequence of prenatal metal and metalloid exposures (37), making it an intermediate on the hypothesized causal pathway. To get a better approximation of the direct effect of prenatal metals/metalloids exposure on the risk of sepsis, we re-fit models explicitly adjusting for all intermediates and their descendants: gestational age, birth weight, histologically-confirmed chorioamnionitis (38), infant sex, preterm premature rupture of membranes (pPROM), and infections during pregnancy. Infections during pregnancy was a composite variable including diagnoses of a urinary tract, vaginal, or group B streptococcus infection, or any antibiotic treatment.

Finally, since animal and epidemiologic studies of in utero metal and metalloid exposures have reported sex-specific associations with immunological endpoints (17, 39), we fit models stratified by infant sex. To formally test for effect measure modification between females and males, we computed two-sample z-test p-values and considered any <0.20 to be significant (40).

Results

Population characteristics

Mothers in our study sample were predominantly over the age of 30 years (56%), non-Hispanic white (61.7%), and not on Medicaid (68.0%) (Table 1). Cord tissue concentrations of several trace elements, notably arsenic, mercury, lead, and selenium, appeared to increase with maternal age. Differences were also observable by maternal race/ethnicity and geographic region. Hispanic mothers gave birth to infants with particularly high levels of
arsenic and lead in their umbilical cords. Mothers who reported smoking during pregnancy gave birth to infants with nearly twice the amounts of cadmium in cord tissue than mothers who reported not smoking. Median concentrations of arsenic, mercury, and lead from New England-born infants were approximately double those born elsewhere. There were slightly more male infants than female (55.4/44.6%). In total, 81 (30.1%, Table 1) infants developed bacterial sepsis during the median follow-up time of 2 weeks (range: 1–4 weeks) and only 3 (1.1%, data not shown) infants were lost to follow-up because they were transferred to a non-ELGAN affiliated hospital. With respect to the trace elements, umbilical cord concentrations were mostly positively correlated with one another with the highest Spearman’s rank correlation coefficient observed for arsenic and mercury ($\rho = 0.5$) (Figure 1).

**Individual trace element models**

We observed no violations of the proportional hazards assumption for any of the 8 trace elements assessed. None of the models for individual trace elements, as measured in umbilical cord tissue, produced significant relationships or linear trends with overall sepsis incidence (Figure 2). However, hazard ratios for increasing quartiles of cadmium were markedly elevated relative to the other trace metals and showed some evidence of a non-monotonic dose-response with the risk of bacterial sepsis. Compared to the lowest quartile, the adjusted hazard ratios for cadmium concentrations in the second, third, and fourth (highest) quartiles were 1.13 (95% CI: 0.51–2.54), 1.94 (95% CI: 0.87–4.32), and 1.88 (95% CI: 0.90–3.93) (Supplementary Table 1). The hazard ratios for some of the other trace elements were noticeably below the null, including the highest quartiles of arsenic, copper, and selenium (Figure 2; Supplementary Table 1).

When focusing on late-onset sepsis, associations were generally consistent with associations for any sepsis (Supplementary Table 2), albeit less precise. Adjusting for hypothesized intermediates also did not materially change the results (Supplementary Table 3).

In sex-stratified models, associations between cadmium and sepsis were strongest for female infants (Supplementary Table 4). The adjusted hazard ratio comparing the highest to the lowest quartile was 4.54 (95% CI: 0.86–23.89) among females, but only 1.19 (95% CI: 0.35–4.03) among males. However, p-values for effect measure modification for cadmium and all other trace elements were above 0.20 (Supplementary Table 4).

**Trace element mixture models**

The quantile g-computation model containing all 8 trace elements estimated an adjusted hazard ratio of 0.92 (95% CI: 0.68–1.25, Figure 3). Of the elements assigned positive weights, cadmium had the largest contribution (Figure 4). Of the elements assigned negative weights, arsenic, copper, and selenium, had the largest contributions. The remaining trace elements were weighted close to zero, indicating negligible associations with incident sepsis. Separating the trace elements into two classes revealed opposing directions of associations: the mixture of non-essential trace elements was associated with an 19% (HR: 1.19, 95% CI: 0.92–1.55) higher risk of sepsis whereas the mixture of essential trace elements was associated with a 23% (HR: 0.77, 95% CI: 0.57–1.06) lower risk (Figure 3).
In quantile g-computation models of late-onset sepsis, the hazard ratios for mixtures of all trace elements and the essential trace elements were similar to estimates for any sepsis, but the hazard ratio for the mixture of non-essential elements was attenuated toward the null (Supplementary Table 5).

In models that were additionally adjusted for hypothesized intermediates (i.e., gestational age, birth weight, chorioamnionitis, infant sex, pPROM, and infections during pregnancy), hazard ratios for the three sets of trace element mixtures with incident sepsis were comparable to those observed in the primary analyses (Supplementary Table 6).

Finally, mixture models stratified by infant sex did not suggest any effect measure modification with p-values >0.20 and similar hazard ratios observed among males and females (Supplementary Table 7).

Discussion

Neonatal bacterial sepsis is a leading cause of infant morbidity and mortality, resulting in a significant healthcare burden (1, 41). Survivors often face neurological sequelae such as cerebral palsy and impaired motor function that have a persistent impact on quality of life (42). Although research suggests that the contribution of environmental factors to sepsis is just as strong as genetics (43), there have been no epidemiologic studies to date on the role of environmental chemical exposures. In this prospective cohort of infants born extremely preterm, we found suggestive evidence that higher levels of prenatal exposure to cadmium were associated with an increased risk of developing bacterial sepsis during the first 28 days of life. However, the results from mixture analyses indicated that joint prenatal exposures to arsenic, copper, and selenium may have a mitigating effect.

Cadmium is found in tobacco smoke and is a contaminant of certain foods (e.g., cereals, bread, and leafy vegetables) (44). Although smoking during pregnancy was associated with notably higher levels of cadmium in umbilical cord tissue, the low prevalence in ELGAN (9.3%) makes it likely that diet was the predominant source of cadmium exposure for this population (45). During pregnancy, cadmium is partially sequestered in the placenta but does cross the placental barrier to reach the developing fetus (46). We found that nearly all umbilical cord tissue samples (98.1%) had detectable levels of cadmium and the median concentration was 1.33 ng/g, a value in line with other studies examining trace elements in cord tissue (47, 48). Prenatal exposure to cadmium has been tied to immunosuppressive endpoints in both animal and epidemiologic studies, including lower levels of T helper memory and natural killer cells (17, 49), potentially increasing susceptibility to infections. There is also evidence to suggest the immunosuppressive effects of prenatal cadmium exposure differ by fetal sex. In mice exposed to environmentally-relevant levels of cadmium in utero and later immunized with the streptococcal vaccine, females exhibited a greater reduction in vaccine-induced antibody production and in the number of antibody secreting cells as compared to males (49). In our study, stratified models revealed stronger associations between prenatal cadmium exposure and incident sepsis in female infants, highlighting the potential for sex-based differences in the immunotoxic effects related to prenatal cadmium exposure.
While our results suggested greater prenatal exposure to cadmium increases the risk of neonatal bacterial sepsis, copper, selenium, and arsenic seemed to do the opposite. Although we anticipated that copper and selenium, as essential nutrients, would confer protection, the results for arsenic were contrary to our hypothesis. Arsenic is a non-essential metalloid and prenatal exposures have previously been linked to immunosuppression and inflammation (50, 51). However, it is possible that arsenic exposure could reduce the risk of sepsis, as it was historically used as a treatment for syphilis due to its antimicrobial properties (52). In 1909, Paul Ehrlich and Sahachiro Hata discovered that the organic arsenic compound arsphenamine (later sold under the brand name Salvarsan) was remarkably effective at killing the parasitic bacterium Treponema pallidum without overtly harming the host (53). One explanation for our findings is that the infants with the highest cord tissue arsenic concentrations had predominantly been exposed to organic arsenic compounds (e.g., arsenobetaine, arsenosugars, arsenolipids), which are frequently found in seafood and are believed to be far less toxic than inorganic forms (54). Since only total arsenic content was measured for this study, we were unable to consider exposure to different arsenic species. It is also plausible that the inverse association observed for arsenic is an artifact of uncontrolled negative confounding by seafood intake (55). In addition to organic arsenicals, seafood, including fish, shellfish, and seaweed, contains nutrients like omega-3 fatty acids that have been shown in an in vitro study to induce anti-inflammatory effects relevant for the prevention of neonatal sepsis (56). The ELGAN cohort did not assess prenatal dietary intakes so we were unable to examine seafood consumption during pregnancy as a source of arsenic and/or mercury exposure. However, the strongest element-by-element correlation we observed was between arsenic and mercury. Methylmercury, in particular, is a common seafood contaminant and, like arsenic, requires more complex sample preparation and speciation analysis techniques to determine specific biomarker concentrations (57). Given the very small sample masses of umbilical cord tissue available, speciation analysis was not feasible for this study. Future cohort studies examining environmental trace element exposures should consider collecting prospective dietary data and sufficient biological samples for performing speciation analyses to better disentangle beneficial and harmful effects.

The apparent protective associations observed for copper and selenium in this study are noteworthy and raise the possibility of dietary interventions as preventive or therapeutic strategies. Similar to the issue of seafood consumption during pregnancy, the findings for copper and selenium may be an artifact of uncontrolled confounding by prenatal supplement use, for which detailed information was not collected in ELGAN. Still, experimental data and the biological role of nutritionally essential trace elements provide a rationale for further investigation in the context of neonatal bacterial sepsis prevention (58–60). In a randomized trial of very low birth weight and preterm infants (<1500 g and <32 weeks’, n = 90), supplementation of 10 μg selenium per day for 28 days significantly reduced the risk of culture-proven sepsis by 13.3% (60). Both copper and selenium are cofactors for antioxidant enzymes, which suppress the formation of free radicals that, if left unchecked, can lead to cellular and tissue damage (61). Specifically, glutathione peroxidase depends on selenium (62), whereas superoxide dismutase, in forms SOD1 and SOD3, requires copper (in addition to zinc) (63). In the United States, it is estimated that <10% of pregnant women have...
inadequate copper and selenium intakes, but virtually no one exceeds the respective tolerable upper intake levels (64), indicating there may be an opportunity to increase amounts present in prenatal dietary supplements. Another potential avenue could be including selenium in parenteral nutrition formulations used by preterm infants. In spite of a recommendation by the American Society for Parenteral and Enteral Nutrition, selenium is not a component in the only multi-trace element product currently marketed in North America for neonates (Multitrace-4 Neonatal by American Regent) (65). Investigating supplementation, either during pregnancy or after delivery, for the prevention of neonatal sepsis is a key area for future research (66), and could be expanded to examine efficacy in sub-groups highly exposed to environmental contaminants.

Although there is some evidence for adverse immunological effects of lead (20), mercury (18), manganese (67), and zinc at certain levels (24), associations for these trace elements appeared null in this analysis. This could be due to the particular exposure ranges represented in ELGAN. For instance, compared to umbilical cord tissue concentrations reported in the literature (47, 48), we found lower concentrations of lead and higher concentrations of zinc, suggesting low levels of exposure to lead pollution and adequate zinc nutriture within the ELGAN cohort. Another important consideration is the relatively small sample size, which may have hampered our ability to detect modest associations and left us underpowered to conduct some proposed analyses. Notably, we were unable to investigate early-onset cases of sepsis independently. In models restricted to late-onset cases, the association for the non-essential trace elements mixture was near null. While this finding may indirectly support in utero exposures to non-essential trace elements as a more salient risk factor for early as opposed to late-onset sepsis, validation in a larger cohort is needed. Alternatively, we acknowledge that although many non-essential trace elements show evidence for immunotoxicity, it may be that neonatal bacterial sepsis is not a particularly sensitive outcome to exposures in the context of extremely preterm infants for whom other factors (e.g., in utero infection status, immune system immaturity, and the need for invasive medical procedures) likely play an outsized role in determining risk.

Despite limitations of this study, notably its small sample size, lack of important dietary information, and inability to perform trace element speciation, it is the first to investigate environmental chemical exposures as a risk factor for neonatal bacterial sepsis. Beyond novelty, there are notable analytic strengths, including the prospective study design, classification of sepsis status through blood culture, and assessment of multiple trace element exposures through biomarkers. Sepsis status was defined by the presence of bacteria in the (usually sterile) bloodstream, indicating bacteremia. At first glance, this definition may seem incomplete because it does not include measures of systemic inflammation. However, markers like C-reactive protein (CRP) and white blood cell counts are generally uninformative in neonates (68). By contrast, blood culture is considered the “gold standard” for diagnosing bacterial sepsis in neonates and was performed serially in ELGAN, improving the internal validity of this study (68). In addition, the use of umbilical cord tissue as a biospecimen to measure prenatal exposure to metallic and metalloid trace elements is a unique advantage. Although most studies rely on maternal samples such as blood or urine collected during pregnancy, maternal biomarkers may not reflect fetal exposures due to toxicokinetics. At present, the literature on trace element content within umbilical
cord tissue is sparse and limited data suggest that, compared to umbilical cord blood, tissue is acceptable for assessing \textit{in utero} exposures to mercury and selenium, whereas it may be the inferior matrix for cadmium, lead, copper, and zinc, and there are no data on arsenic or manganese either way (48). In addition, there is the possibility that the location of the tissue sampled affects the amount of trace element present; given that the tissue assessed in this study was obtained from the portion of umbilical cord closest to the abdominal wall, we believe our study appropriately characterized prenatal trace element exposures of mercury and selenium but acknowledge that instead measuring concentrations of the other trace elements in cord blood would have likely enhanced study validity. An additional strength of this study was the availability of rich covariate information, including sociodemographic factors and pregnancy complications that are known risk factors for neonatal sepsis. We leveraged these data to control for potential confounders in primary analyses and hypothesized intermediates in supplemental analyses. The results from both were nearly identical, indicating that, under certain assumptions (69), prenatal metal and metalloid exposures likely influence the risk of neonatal bacterial sepsis directly, rather than through pregnancy complications or infant characteristics.

In summary, we identified a positive association for prenatal cadmium exposure with incident bacterial sepsis in a cohort of extremely preterm infants, which was consistent when cadmium was examined in isolation and within the context of joint exposures to other metals and metalloids. This association was more pronounced among female infants, suggesting cadmium-associated immunotoxicity may differ by sex. It also did not appear to be mediated through pregnancy complications or infant characteristics, indicating cadmium may directly affect the developing immune system. Interestingly, higher levels of \textit{in utero} exposures to arsenic, copper, and selenium seemed to mitigate the risk associated with high cadmium exposures. Future studies are needed to understand whether exposure to specific arsenic species are contributing to this unexpected finding, whether the results could be explained by negative confounding by seafood consumption during pregnancy, and whether copper and selenium supplementation are viable strategies for reducing rates of bacterial sepsis in high-risk infants.

\textbf{Supplementary Material}

Refer to Web version on PubMed Central for supplementary material.

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Data availability statement

The data that support the findings of this study are available upon reasonable request from the corresponding author, RCF. The data are not available publicly due to their containing information that could compromise the privacy of the ELGAN study participants.

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FIGURE 1.
Heatmap displaying Spearman’s rank correlation coefficients between all possible pairs of the trace elements, as measured in ELGAN umbilical cord tissue samples ($n = 269$).
FIGURE 2.
Forest plots showing adjusted* hazard ratios with 95% confidence intervals for neonatal bacterial sepsis in the first 28 days of life by trace element quartiles. $p_{\text{trend}}$ for linear associations with sepsis calculated by modeling the median trace element concentration in each quartile as a continuous variable. *Adjusted for geographic region, maternal race/ethnicity, age, educational attainment, Medicaid coverage, pre-pregnancy BMI, and smoking status during pregnancy.
FIGURE 3.
Forest plot showing adjusted* hazard ratios with 95% confidence intervals for neonatal bacterial sepsis in the first 28 days of life by trace element mixture. *Adjusted for geographic region, maternal race/ethnicity, age, educational attainment, Medicaid coverage, pre-pregnancy BMI, and smoking status during pregnancy.
FIGURE 4.
Bar plot showing the weights assigned to each trace element composing the overall mixture as estimated by the quantile g-computation model for neonatal bacterial sepsis in the first 28 days of life after adjustment for geographic region, maternal race/ethnicity, age, educational attainment, Medicaid coverage, pre-pregnancy BMI, and smoking status during pregnancy. The size of the bar corresponds to the proportional contribution of a given trace element to the association with sepsis in a particular direction (negative or positive). Because the size of the bars should only be compared relative to others in the same direction, the shading reflects the absolute contribution of the individual trace element, regardless of directionality.
TABLE 1

Median umbilical cord tissue trace element concentrations, overall and stratified by participant characteristics in the ELGAN study ($n = 269$).

|                          | Non-essential trace elements (ng/g) | Essential trace elements (μg/g) |
|--------------------------|------------------------------------|----------------------------------|
|                          | Arsenic | Cadmium | Mercury | Lead | Copper | Manganese | Selenium | Zinc |
| Overall                  | 269 (100.0) | 4.72 (3.32, 7.61) | 1.33 (0.76, 3.11) | 7.74 (3.66, 17.65) | 16.60 (9.30, 31.60) | 3.48 (2.97, 4.32) | 0.34 (0.30, 0.42) | 0.86 (0.77, 0.97) | 60.30 (52.23, 71.70) |
| Maternal age             |         |         |         |       |        |           |          |      |
| ≤25 years                | 71 (26.4) | 3.80 (2.46, 5.19) | 0.99 (0.68, 2.47) | 5.62 (2.29, 10.84) | 12.60 (8.15, 25.05) | 3.26 (2.81, 4.11) | 0.34 (0.30, 0.42) | 0.83 (0.74, 0.96) | 56.30 (46.65, 63.10) |
| >25–≤30 years            | 74 (27.5) | 5.18 (3.32, 7.14) | 1.28 (0.75, 2.41) | 6.00 (3.14, 13.68) | 13.25 (8.77, 21.92) | 3.62 (3.08, 4.39) | 0.35 (0.28, 0.42) | 0.86 (0.81, 0.98) | 61.90 (53.83, 74.07) |
| >30–≤35 years            | 62 (23.0) | 4.75 (3.39, 9.22) | 1.35 (0.77, 3.01) | 10.75 (5.27, 17.84) | 19.30 (11.20, 31.90) | 3.79 (3.13, 4.36) | 0.35 (0.30, 0.42) | 0.86 (0.77, 0.95) | 61.30 (54.27, 72.80) |
| >35 years                | 62 (23.0) | 6.36 (4.30, 9.80) | 1.79 (1.03, 3.73) | 15.54 (6.14, 32.62) | 21.85 (13.65, 43.50) | 3.40 (3.01, 4.58) | 0.33 (0.30, 0.42) | 0.87 (0.80, 0.96) | 61.85 (54.60, 74.93) |
| Maternal race/ethnicity  |         |         |         |       |        |           |          |      |
| Hispanic                 | 20 (7.4) | 6.00 (4.38, 8.79) | 1.21 (0.82, 2.32) | 11.38 (4.58, 18.26) | 21.70 (12.60, 34.29) | 3.62 (3.35, 4.94) | 0.32 (0.28, 0.37) | 0.81 (0.73, 0.90) | 59.13 (53.20, 66.77) |
| Non-Hispanic black       | 73 (27.1) | 3.80 (2.50, 6.75) | 0.88 (0.58, 2.21) | 6.70 (4.34, 13.69) | 11.30 (7.40, 19.50) | 3.50 (2.87, 4.47) | 0.34 (0.28, 0.43) | 0.91 (0.77, 0.99) | 55.30 (47.60, 62.99) |
| Non-Hispanic other       | 10 (3.7) | 4.14 (3.32, 5.47) | 1.55 (1.02, 3.66) | 11.06 (3.48, 14.79) | 16.00 (12.62, 26.75) | 3.23 (3.02, 3.67) | 0.44 (0.39, 0.52) | 0.80 (0.71, 0.86) | 60.05 (47.60, 69.70) |
| Non-Hispanic white       | 166 (61.7) | 4.97 (3.62, 8.04) | 1.54 (0.86, 3.75) | 8.29 (3.40, 18.95) | 18.65 (10.12, 35.10) | 3.47 (2.97, 4.24) | 0.34 (0.30, 0.41) | 0.86 (0.78, 0.96) | 63.05 (54.94, 77.77) |
| Maternal educational attainment |         |         |         |       |        |           |          |      |
| Less than high school    | 34 (12.6) | 3.93 (2.45, 5.45) | 1.65 (0.79, 3.34) | 6.17 (2.99, 14.37) | 14.60 (11.93, 32.04) | 3.42 (3.03, 4.18) | 0.37 (0.30, 0.40) | 0.84 (0.75, 0.92) | 52.90 (42.42, 63.70) |
| High school diploma      | 124 (46.1) | 4.37 (2.81, 6.51) | 1.19 (0.76, 2.73) | 6.58 (3.34, 12.92) | 15.00 (8.17, 27.75) | 3.45 (2.89, 4.23) | 0.33 (0.30, 0.42) | 0.85 (0.76, 0.98) | 57.80 (48.90, 66.72) |
| College                  | 111 (41.3) | 5.70 (4.01, 9.85) | 1.36 (0.75, 3.27) | 11.19 (5.31, 26.43) | 19.40 (10.90, 33.25) | 3.58 (3.00, 4.49) | 0.35 (0.30, 0.41) | 0.88 (0.80, 0.96) | 64.70 (56.25, 75.90) |
| Maternal medicaid coverage |         |         |         |       |        |           |          |      |
| No                       | 183 (68.0) | 5.50 (3.64, 8.93) | 1.35 (0.78, 3.00) | 10.41 (4.94, 21.55) | 18.20 (10.30, 33.99) | 3.54 (2.98, 4.24) | 0.35 (0.30, 0.42) | 0.87 (0.79, 0.97) | 62.80 (55.09, 75.50) |
| Yes                      | 86 (32.0) | 3.99 (2.47, 5.40) | 1.19 (0.72, 3.42) | 5.57 (2.62, 10.46) | 12.60 (8.12, 22.28) | 3.42 (2.94, 4.55) | 0.33 (0.29, 0.42) | 0.82 (0.73, 0.94) | 54.80 (46.38, 62.53) |
| Maternal smoking during pregnancy |         |         |         |       |        |           |          |      |
|                          | n (%) | Non-essential trace elements (ng/g) | Essential trace elements (μg/g) |
|--------------------------|-------|------------------------------------|-------------------------------|
|                          |       | Arsenic, Cd., 7.89                | Copper, 4.16, 4.46             |
|                          |       | Cadmium, 1.27, 2.86               | Manganese, 0.20, 0.42          |
|                          |       | Mercury, 8.50, 18.04              | Selenium, 0.86, 0.97           |
|                          |       | Lead, 16.60, 30.85               | Zinc, 61.10, 52.68             |
| No                       | 244 (90.7) | 4.71 (3.32, 7.89)               | 3.54 (2.99, 4.46)             |
|                          |       | 1.27 (0.75, 2.86)               | 0.34 (0.30, 0.42)             |
|                          |       | 8.50 (3.83, 18.04)              | 0.86 (0.79, 0.97)             |
|                          | Yes   | 25 (9.3)                        | 14.70 (11.20, 35.90)          |
|                          |       | 4.74 (3.38, 5.50)               | 2.82 (2.89, 3.87)             |
|                          |       | 2.45 (1.28, 14.89)              | 0.37 (0.29, 0.40)             |
|                          |       | 5.52 (2.78, 10.86)              | 0.80 (0.73, 0.96)             |
|                          |       | 61.10 (52.68, 73.32)            | 54.20 (51.00, 60.10)          |
| Maternal pre-pregnancy body mass index |       |                                   |                               |
| Normal                   | 167 (62.1) | 4.69 (3.42, 7.52)               | 17.20 (9.65, 32.70)           |
|                          |       | 1.34 (0.78, 2.91)               | 3.41 (2.88, 4.13)             |
|                          |       | 7.97 (3.33, 19.36)              | 0.34 (0.30, 0.41)             |
|                          | Obese | 58 (21.6)                       | 16.20 (10.08, 29.80)          |
|                          |       | 4.44 (2.80, 6.46)               | 3.88 (3.21, 4.74)             |
|                          |       | 1.15 (0.69, 3.49)               | 0.33 (0.28, 0.41)             |
|                          |       | 6.63 (4.06, 11.96)              | 0.83 (0.73, 0.91)             |
|                          | Overweight | 44 (16.4)                       | 11.90 (8.97, 21.33)          |
|                          |       | 5.78 (3.56, 9.81)               | 3.46 (3.12, 4.23)             |
|                          |       | 1.42 (0.68, 3.39)               | 0.35 (0.30, 0.46)             |
|                          |       | 10.87 (4.15, 19.56)             | 0.86 (0.80, 0.97)             |
|                          |       | 63.70 (54.83, 72.83)            | 63.70 (54.83, 72.83)          |
| Geographic region        |       |                                   |                               |
| Midwest                  | 68 (25.3) | 4.65 (3.26, 5.82)               | 10.72 (6.68, 17.45)           |
|                          |       | 1.60 (0.72, 4.38)               | 3.59 (2.91, 4.36)             |
|                          |       | 4.20 (2.22, 9.57)               | 0.32 (0.29, 0.41)             |
|                          | New England | 121 (45.0)                     | 23.90 (17.00, 43.40)          |
|                          |       | 5.96 (4.20, 9.90)               | 3.54 (3.07, 4.42)             |
|                          |       | 1.55 (0.85, 3.31)               | 0.36 (0.31, 0.43)             |
|                          |       | 15.05 (8.13, 32.04)             | 0.87 (0.80, 0.96)             |
|                          | North Carolina | 80 (29.7)                       | 11.15 (7.18, 19.27)          |
|                          |       | 3.32 (2.32, 5.20)               | 3.30 (2.89, 4.01)             |
|                          |       | 1.03 (0.70, 1.96)               | 0.32 (0.28, 0.42)             |
|                          |       | 6.00 (3.26, 8.73)               | 0.77 (0.70, 0.85)             |
|                          |       | 65.30 (56.22, 79.35)            | 50.45 (44.27, 58.02)          |
| Infant sex               |       |                                   |                               |
| Male                     | 149 (55.4) | 4.73 (3.32, 8.11)               | 18.20 (10.83, 34.00)          |
|                          |       | 1.35 (0.81, 3.31)               | 3.59 (3.02, 4.53)             |
|                          |       | 8.36 (4.02, 20.23)              | 0.34 (0.30, 0.41)             |
|                          | Female | 120 (44.6)                       | 14.40 (8.82, 25.83)           |
|                          |       | 4.70 (3.32, 6.95)               | 3.39 (2.88, 4.09)             |
|                          |       | 1.25 (0.73, 2.49)               | 0.34 (0.30, 0.42)             |
|                          |       | 7.32 (3.43, 15.06)              | 0.86 (0.77, 0.97)             |
|                          |       | 61.50 (53.95, 73.30)            | 57.94 (50.94, 69.90)          |
| Neonatal bacterial Sepsis |       |                                   |                               |
| No                       | 188 (69.9) | 5.05 (3.42, 8.64)               | 16.60 (9.97, 31.38)           |
|                          |       | 1.22 (0.74, 2.59)               | 3.54 (3.02, 4.46)             |
|                          |       | 8.56 (3.74, 20.22)              | 0.34 (0.30, 0.42)             |
|                          | Yes   | 81 (30.1)                       | 16.30 (8.70, 32.90)           |
|                          |       | 4.51 (3.07, 5.70)               | 3.39 (2.86, 3.99)             |
|                          |       | 1.84 (0.86, 3.98)               | 0.34 (0.29, 0.43)             |
| Early-onset              | 19 (7.1) | 5.00 (3.71, 7.13)               | 16.30 (9.45, 31.40)           |
|                          |       | 2.15 (1.41, 4.80)               | 3.83 (2.69, 4.19)             |
|                          | Late-onset | 62 (23.0)                      | 16.17 (7.92, 32.94)          |
|                          |       | 4.14 (2.84, 5.69)               | 3.33 (2.89, 3.85)             |
|                          |       | 1.77 (0.85, 3.93)               | 0.34 (0.29, 0.41)             |
|                          |       | 5.46 (2.84, 11.32)              | 0.84 (0.74, 0.94)             |
|                          |       | 61.04 (51.75, 78.90)            | 61.04 (51.75, 78.90)          |

Values inside the parenthesis represent the 25th and 75th percentiles.