Association of Ketamine With Psychiatric Symptoms and Implications for Its Therapeutic Use and for Understanding Schizophrenia: A Systematic Review and Meta-analysis

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Abstract

Importance Ketamine hydrochloride is increasingly used to treat depression and other psychiatric disorders but can induce schizophrenia-like or psychotomimetic symptoms. Despite this risk, the consistency and magnitude of symptoms induced by ketamine or what factors are associated with these symptoms remain unknown.

Objective To conduct a meta-analysis of the psychopathological outcomes associated with ketamine in healthy volunteers and patients with schizophrenia and the experimental factors associated with these outcomes.

Data Sources MEDLINE, Embase, and PsychINFO databases were searched for within-participant, placebo-controlled studies reporting symptoms using the Brief Psychiatric Rating Scale (BPRS) or the Positive and Negative Syndrome Scale (PANSS) in response to an acute ketamine challenge in healthy participants or patients with schizophrenia.

Study Selection Of 8464 citations retrieved, 36 studies involving healthy participants were included. Inclusion criteria were studies (1) including healthy participants; (2) reporting symptoms occurring in response to acute administration of subanesthetic doses of ketamine (racemic ketamine, s-ketamine, r-ketamine) intravenously; (3) containing a placebo condition with a within-subject, crossover design; (4) measuring total positive or negative symptoms using BPRS or PANSS; and (5) providing data allowing the estimation of the mean difference and deviation between the ketamine and placebo condition.

Data Extraction and Synthesis Two independent investigators extracted study-level data for a random-effects meta-analysis. Total, positive, and negative BPRS and PANSS scores were extracted. Subgroup analyses were conducted examining the effects of blinding status, ketamine preparation, infusion method, and time between ketamine and placebo conditions. The Meta-analysis of Observational Studies in Epidemiology (MOOSE) and Preferred Reporting Items for Systematic Reviews and Meta-analyses (PRISMA) guidelines were followed.

Main Outcomes and Measures Standardized mean differences (SMDs) were used as effect sizes for individual studies. Standardized mean differences between ketamine and placebo conditions were calculated for total, positive, and negative BPRS and PANSS scores.

Results The overall sample included 725 healthy volunteers (mean [SD] age, 28.3 [3.6] years; 533 [73.6%] male) exposed to the ketamine and placebo conditions. Racemic ketamine or S-ketamine was associated with a statistically significant increase in transient psychopathology in healthy

Key Points

Question What psychopathological outcomes are associated with ketamine hydrochloride in healthy volunteers and patients with schizophrenia, and what factors are associated with these outcomes?

Findings This meta-analysis of 36 studies, including 725 unique healthy participants, found that the acute administration of ketamine relative to placebo was associated with a meaningful increase in positive and negative symptoms of psychosis in healthy volunteers and patients with schizophrenia. This association was greater for positive compared with negative symptoms and when a bolus was given with the infusion relative to an infusion alone.

Meaning These findings suggest that ketamine is associated with psychosis-like symptoms in healthy volunteers and that the bolus administration of ketamine should be avoided when it is used in therapeutic contexts.

Supplemental content

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Abstract (continued)

participants for total (SMD = 1.50 [95% CI, 1.23 -1.77]; \( P < .001 \)), positive (SMD = 1.55 [95% CI, 1.29 -1.81]; \( P < .001 \)), and negative (SMD = 1.16 [95% CI, 0.96 -1.35]; \( P < .001 \)) symptom ratings relative to the placebo condition. The effect size for this association was significantly greater for positive than negative symptoms of psychosis (estimate, 0.36 [95% CI, 0.12 -0.61]; \( P = .004 \)). There was significant inconsistency in outcomes between studies (\( I^2 \) range, 77%-83%). Bolus followed by constant infusion increased ketamine’s association with positive symptoms relative to infusion alone (effect size, 1.63 [95% CI, 1.36 -1.90] vs 0.84 [95% CI, 0.35 -1.33]; \( P = .006 \)). Single-day study design increased ketamine’s ability to generate total symptoms (effect size, 2.29 [95% CI, 1.69 -2.89] vs 1.39 [95% CI, 1.12 -1.66]; \( P = .007 \)), but age and sex did not moderate outcomes. Insufficient studies were available for meta-analysis of studies in schizophrenia. Of these studies, 2 found a statistically significant increase in symptoms with ketamine administration in total and positive symptoms. Only 1 study found an increase in negative symptom severity with ketamine.

CONCLUSIONS AND RELEVANCE This study found that acute ketamine administration was associated with schizophrenia-like or psychotomimetic symptoms with large effect sizes, but there was a greater increase in positive than negative symptoms and when a bolus was used. These findings suggest that bolus doses should be avoided in the therapeutic use of ketamine to minimize the risk of inducing transient positive (psychotic) symptoms.

Introduction

Ketamine hydrochloride was first synthesized in 1962.\(^1\) It is a phencyclidine derivative that acts on the glutamate system by antagonizing N-methyl-D-aspartate (NMDA) receptors.\(^1\) Ketamine has been used to model the symptoms of schizophrenia and is used in the treatment of severe depression and pain management\(^2\) as well as being used recreationally. Misuse can be hazardous, leading to drug addiction.

In the 1960s, NMDA antagonists, such as ketamine, were identified as inducing clinical symptoms similar to those seen in schizophrenia, more so than other psychotomimetics used in past drug models of psychosis.\(^3,4\) In particular, in addition to inducing positive symptoms, such as perceptual changes and delusions, ketamine induces negative symptoms, such as blunted affect and emotional withdrawal.\(^5\) Many studies have been conducted to investigate its effect on healthy people, but the methods vary greatly, and the observed behavioral responses differ.

Despite the recognition that ketamine can induce transient schizophrenia-like symptoms,\(^6\) the consistency and magnitude of its effect on positive and negative symptoms remains unclear. Moreover, it is unclear how blinding status, ketamine preparation, infusion method, and time between the ketamine and placebo conditions are associated with the generation of symptoms.

The development of ketamine and its derivatives as antidepressants\(^6,7\) means that determining the extent to which ketamine induces schizophrenia-like or psychotomimetic symptoms and what factors are associated with this outcome is particularly timely in order to understand and minimize the risks of adverse events associated with the therapeutic use of ketamine. We also aimed to evaluate outcomes in patients with schizophrenia to determine whether they are more sensitive to ketamine.

We therefore conducted a systematic review and meta-analysis of the association of ketamine with positive, negative, and total psychopathological outcomes in healthy volunteers and patients with schizophrenia. Many studies use ketamine to inform understanding of the mechanisms underlying schizophrenia. This specific use of ketamine is the main focus of our review, but we also use the findings to inform understanding of other uses of ketamine.
Methods

Selection Procedures
A meta-analysis was performed according to the Meta-analysis of Observational Studies in Epidemiology (MOOSE) and Preferred Reporting Items for Systematic Reviews and Meta-analyses (PRISMA) frameworks. Three authors (K.B., G.H., and F.B.) independently searched MEDLINE (from 1946 to February 3, 2020), Embase (from 1974 to February 3, 2020), and PsychINFO (from 1806 to January 27, 2020). The following keywords were used: (Ketamine) and (psycho* NOT psychotherapy or schiz* or BPRS or brief psychiatric rating scale or PANSS or positive and negative syndrome scale or positive symp* or negative symp*). Meta-analyses and systematic and narrative review articles were hand-searched for additional reports. Abstracts were screened, and the full texts of suitable studies were obtained. If full texts were not available, authors were contacted and full content was requested. Authors were also contacted when Brief Psychiatric Rating Scale (BPRS) or Positive and Negative Syndrome Scale (PANSS) subscales (total, negative, or positive) were missing or if the individual items included in the positive or negative scales were not reported. Three authors (K.B., G.H., and F.B.) selected the final studies included in the meta-analysis based on the following criteria.

Selection Criteria for the Meta-analysis of Ketamine's Effects in Healthy Volunteers
Inclusion criteria were studies (1) including healthy participants, (2) reporting symptoms occurring in response to acute administration of subanesthetic doses of ketamine (racemic ketamine, s-ketamine, or r-ketamine) intravenously, (3) containing a placebo condition with a within-participant, crossover design, (4) measuring total positive or negative symptoms using the BPRS or PANSS, and (5) providing data allowing the estimation of the mean difference and deviation between the ketamine and placebo condition. We used the PANSS and BPRS scales as the measures of symptom severity because they are well validated, standardized assessments of psychopathology used in both healthy participants and patients with schizophrenia. They assess the same symptom dimensions and are commonly combined in meta-analyses. We included all versions of the total BPRS because often the version was not specified. All versions measure the same rating items, but some include more items than others. However, all included studies are within-person studies, and so this should not affect the analysis. Exclusion criteria consisted of 1 or more of the following factors: (1) no placebo condition, (2) no report of any total, negative, or positive scores (see the following sections for more details), (3) absence of measures in either the ketamine or the placebo condition, (4) no report of original data, (5) no data provided that enabled the standardized mean differences (SMDs) to be calculated (such as the SD or the standard error of the mean), (6) no more than 2 participants in each group, and/or (7) concurrent administration of other pharmacological compounds in addition to ketamine.

Selection Criteria for the Meta-analysis of Ketamine's Effects in Schizophrenia
The selection criteria for studies investigating the effect of ketamine in patients with schizophrenia were the same as the criteria for healthy volunteers. The only additional criterion was for participants to have a Diagnostic and Statistical Manual of Mental Disorders or International Statistical Classification of Diseases and Related Health Problems, Tenth Revision diagnosis of schizophrenia or schizoaffective disorder.

Additional Symptom Subdomain Inclusion Criteria for Both Meta-analyses
Studies used different combinations of symptom items in their positive and negative BPRS scores. We included studies in the negative analysis if their BPRS scale included all 3 negative symptom items: blunted affect, emotional withdrawal, and motor retardation. We included studies in the positive analysis if they included more than 3 positive symptom items: conceptual disorganization, hallucinatory behavior, unusual thought content, and suspiciousness. These symptom items correlate most strongly and reliably with validated scales of positive and negative symptoms: the
Scale for the Assessment of Positive Symptoms\textsuperscript{14} and Scale for the Assessment of Negative Symptoms,\textsuperscript{15} respectively. If the symptom items included in the scale were not documented, we requested the information from authors (eMethods 1 in the Supplement).

**Recorded Variables**

The primary outcome measures were the effect sizes for total, positive, and negative BPRS and PANSS scores in healthy participants or in patients with schizophrenia for ketamine compared with placebo conditions. Data were extracted from every study for author, year of publication, number of participants, participant age, sex, diagnosis, study design, details of the placebo condition, past or present psychiatric diagnoses among healthy volunteers, recent substance misuse or dependence history, family history of psychosis, major medical or neurological disorder, prior exposure to ketamine, concurrent psychotropic medication use, ketamine preparation, dose and timing of ketamine administration relative to the symptom measures, and mean (SD) measure of symptoms in the ketamine and placebo conditions. Plot digitizer software was used to examine reliability for the data from studies in which data were only available in a plot format.

The highest available ketamine dose was selected if multiple doses were reported. All data sets included in the meta-analysis were independent, and there was no overlap in the participants included in the meta-analyses. The raw data are provided in eTables 1 to 4 in the Supplement.

**Risk of Bias**

Risk of bias was assessed using the Newcastle-Ottawa tool for assessing risk in nonrandomized studies and the Cochrane assessment of risk of bias tool.\textsuperscript{16,17} Scores were calculated by 2 investigators (K.B., G.H.). Studies with scores of at least 7 were considered to have a low risk of bias (eMethods 2-4 in the Supplement).

**Statistical Analysis**

Statistical analyses were conducted using the metafor package, version 1.9-9, with R software, version 3.3.1 (R Project for Statistical Computing). Random-effects models based on restricted maximum likelihood estimation were used in all analyses. Random-effects models were deemed preferable for this analysis owing to substantial between-study differences in study design. Effect sizes or SMDs for individual studies were estimated by calculating the standardized mean change scores. Mean differences in symptom measurements between the ketamine and placebo conditions were used to calculate the standardized mean change score. The 95\% CI of the effect size was also calculated.

The SMD was defined for each study as follows\textsuperscript{18}:

\[
\text{SMD} = \frac{M_{\text{ket}} - M_{\text{sal}}}{\sqrt{(SD_{\text{ket}}^2 + SD_{\text{sal}}^2 - 2rSD_{\text{ket}}SD_{\text{sal}})}}
\]

where \(M_{\text{ket}}\) and \(M_{\text{sal}}\) are the mean scores and \(SD_{\text{ket}}\) and \(SD_{\text{sal}}\) are the SDs for the ketamine and saline (placebo) conditions, respectively, with \(r\) denoting the between-condition correlation for symptom scores under saline and ketamine conditions. The correlation coefficient was set to 0.5 for all studies in our main analysis based on evidence from studies in schizophrenia.\textsuperscript{19,20} However, a sensitivity analysis was performed to evaluate the influence of this assumption on our main results by refitting our model using \(r\) values ranging from 0.1 and 0.7 (eMethods 5 and 6 in the Supplement).

To determine whether ketamine had a greater association with positive or negative symptoms, a multivariate meta-analytic approach was adopted using an unstructured variance-covariance matrix. Because within-study correlations between positive and negative symptom scores are not reported, we estimated the correlation coefficient to be 0.5 based on prior studies.\textsuperscript{20} To investigate the influence of this value on the findings, we conducted sensitivity analyses using correlation coefficients of 0.1 and 0.7 (eMethods 7 and 8 in the Supplement).\textsuperscript{21}
Inconsistency or heterogeneity across studies was assessed using the Cochran Q statistic and \( I^2 \) statistic. An \( I^2 \) statistic of less than 25% was taken to indicate low inconsistency; 25% to 75%, medium inconsistency; and greater than 75%, high inconsistency. The \( I^2 \) statistics were calculated for each subgroup analysis. Leave-one-out sensitivity analyses were also conducted.

Publication bias and selective reporting were assessed using the Egger regression test of the intercept and were represented diagrammatically with funnel plots as recommended by the Cochrane Collaboration (eFigures 1-3 in the Supplement). Trim-and-fill analyses were also conducted.

Secondary subgroup and meta-regression analyses were performed to examine the effects of study design. Specifically, we compared effect size in double-blind vs single-blind or unblinded studies; s-ketamine vs racemic ketamine; bolus followed by constant infusion administration vs infusion alone; and single-day (ketamine and placebo were given on the same day) vs multiple-day (ketamine and placebo given on different days) studies. In addition, we compared the effect size from studies using the BPRS with those using the PANSS to determine whether the method of measuring symptoms was associated with the magnitude of the effect. The statistical significance of subgroup differences was determined by fitting separate random-effects models for each subgroup and then comparing subgroup summary estimates in a fixed-effects model with a Wald-type test. A significance level of \( P < .05 \) (2 tailed) was adopted (see eMethods 9 and 10 in the Supplement for further details).

Results

Retrieved Studies for the Meta-analysis of Healthy Volunteers

A total of 36 studies involving 725 unique participants (mean [SD] age, 28.3 [3.6] years; 533 male [73.6%] and 192 female [26.5%]) were included in the meta-analysis. Figure 1 shows the PRISMA flowchart. The included studies are summarized in the Table, with further details in eTable 5 in the Supplement. Ketamine was administered intravenously in all studies. The search identified 2 additional studies using inhaled administration, but these did not have data available. 

Figure 1. Search Process Summarizing the Review and Exclusion of Studies

- 8464 Studies identified in Embase, MEDLINE, and PsycINFO screened
- 5896 Excluded after title and abstract review because topic not pertinent to this meta-analysis
- 2442 Duplicates removed
- 126 Identified for full-text review or full text sought from authors
- 90 Full texts or abstracts excluded
  - 28 Duplicate data
  - 25 Unable to provide raw data
  - 22 No placebo condition
  - 9 Independent group design
  - 3 No IV administration or <2 participants
  - 2 Incorrect symptom items (BPRS)
  - 1 Additional compound given
- 36 Included in the meta-analysis
  - 25 Examine total symptoms
  - 22 Examine negative symptoms
  - 21 Examine positive symptoms
  - 3 Included in the systematic review of patients with schizophrenia

BPRS indicates Brief Psychiatric Rating Scale; IV, intravenous.
| Source            | Sample size, No. | Age, mean (SD), y | Sex, No. male:female | Blinded Randomized | Placebo condition | Symptom subscales reported | Length of ketamine infusion before symptom assessment, min |
|-------------------|------------------|-------------------|----------------------|--------------------|-------------------|---------------------------|----------------------------------------------------------|
| BPRS Studies      |                  |                   |                      |                    |                   |                           |                                                          |
| Kraguljac et al, 30 2017 | 15               | 24.8 (3.49)       | 10:5                 | No                 | No                | Saline                    | Total, positive (2), and negative                        |
| Kort et al, 43 2017         | 31               | 27.0 (4.3)        | 19:12                | Double             | Yes               | Saline                    | Total                                                   |
| Duncan et al, 36 2001      | 16               | 33.3 (3.1)        | 16:0                 | Double             | Yes               | Saline                    | Total and negative                                       |
| Parwani et al, 37 2005    | 13               | 31.9 (9.6)        | 5:8                  | Double             | Yes               | Saline                    | Total                                                   |
| Rowland et al, 38 2005    | 10               | 24.7 (3.4)        | 10:0                 | Double             | Yes               | Saline                    | Total                                                   |
| Abel et al, 42 2003       | 8                | 28.75             |                      | Double             | Yes               | Saline                    | Total                                                   |
| Anand et al, 44 2000      | 16               | 34.0 (12.0)       | 8:8                  | Double             | Yes               | Saline                    | Positive (1) and negative                               |
| Krystal et al, 35 1998    | 23               | 30.0              | 19:11                | Double             | Yes               | Saline                    | Positive (1) and negative                               |
| Breier et al, 33 1997     | 17               | 30.4 (6.8)        | 15:2                 | Double             | Yes               | Saline                    | Positive (1)                                            |
| van Berkel et al, 40 1998 | 18               | 23.7 (2.4)        | 18:0                 | Double             | Yes               | NR                        | Total                                                   |
| Malhotra et al, 25 1997   | 16               | 27.8 (1.9)        | 12:4                 | Double             | Yes               | Saline                    | Total and negative                                       |
| Krystal et al, 45 1999    | 20               | 28                | 10:10                | Double             | Yes               | Saline                    | Positive (1) and negative                               |
| Krystal et al, 46 2003    | 26               | 29.1 (9)          | 19:7                 | Double             | Yes               | Saline                    | Positive (1) and negative                               |
| Micallef et al, 47 2002   | 8                | 27.0              | 4:4                  | Double             | Yes               | Saline                    | Positive (1) and negative                               |
| Rowland et al, 52 2010    | 9                | 30.8              | 4:5                  | Double             | Yes               | Saline                    | Total                                                   |
| Newcomer et al, 31 1999   | 15               | 21.7 (3.2)        | 15:0                 | Double             | Yes               | Saline                    | Total and positive                                       |
| Stone et al, 33 2011      | 8                | 28.5 (9.9)        | 8:0                  | Double             | Yes               | Saline                    | Total                                                   |
| Boeijinga et al, 48 2007  | 12               | 39.6 (4.8)        | 12:0                 | Double             | Yes               | Saline                    | Total                                                   |
| Abdallah et al, 54 2018   | 14               | NR                | NR                   | Single             | No                | Saline                    | Total and negative                                       |
| Passie et al, 55 2003     | 12               | 26.8 (3.31)       | 12:0                 | Double             | Yes               | Saline                    | Total                                                   |
| Horacek et al, 56 2010    | 20               | 29.9 (5.69)       | 13:7                 | Double             | Yes               | Saline                    | Total                                                   |
| Morgan et al, 57 2011     | 16               | 22.4              | 10:8                 | Double             | No                | Saline                    | Total                                                   |
| PANSS Studies           |                  |                   |                      |                    |                   |                           |                                                          |
| Thiebes et al, 29 2017   | 24               | 25.2 (6.4)        | 24:0                 | Single             | Yes               | Saline                    | Total, positive, negative, and factor scores            |
| Powers et al, 58 2015    | 19               | 27.5              | 10:10                | No                 | No                | Saline                    | Positive and negative                                    |
| Höflisch et al, 49 2015  | 30               | 25.4 (4.58)       | NR                   | Double             | Yes               | Saline                    | Total, positive, and negative                           |
| Nagels et al, 59 2011    | 15               | 27.3 (3.6)        | 15:0                 | Double             | Yes               | Saline                    | Total, positive, and negative                           |
| Driesen et al, 60 2013   | 22               | 29.14 (7.07)      | 14:8                 | No                 | No                | Saline                    | Positive and negative                                    |
| Vernaleken et al, 50 2013 | 10              | 24.4 (3.9)        | 10:0                 | Single             | Yes               | Saline                    | Total, positive, and negative                           |
| Krystal et al, 1 2005    | 27               | 30.96             | 16:11                | Double             | Yes               | Saline                    | Total, positive, negative, and factor score             |
| Krystal et al, 72 2006   | 31               | 28.1 (7.6)        | NR                   | Double             | Yes               | Saline                    | Total, positive, negative, and factor score             |
| Kleinloog et al, 28 2015 | 30               | NR                | 15:15                | Double             | Yes               | Saline                    | Positive and negative                                    |
| D’Souza et al, 51 2012   | 32               | 27.8 (8.42)       | NR                   | Double             | Yes               | Saline                    | Total, positive, and negative                           |
| Grent-‘t-Jong et al, 32 2018 | 14         | 29.0 (9.9)        | 12:2                 | Single             | Yes               | Saline                    | Total, negative, and positive                           |
| D’Souza et al, 34 2018   | 26               | 29.8 (9.56)       | 21:5                 | NR                 | No                | Saline                    | Negative and positive                                    |
| Mathalon et al, 61 2014  | 9                | 29.8 (7.9)        | 5:4                  | Double             | Yes               | Saline                    | Total                                                   |

(continued)
Total Psychopathological Symptoms

Total symptom scores were analyzed using data from 25 studies \cite{25,29,33,36,39,40,43,48,50,52,54,57,59,61} including 491 healthy participants exposed to the ketamine and placebo conditions. Total symptom scores were increased in the ketamine condition compared with the placebo condition (SMD = 1.50 [95% CI, 1.23-1.77]; $P < .001$) (Figure 2). The finding remained statistically significant in all iterations of the leave-one-out analysis (SMD range, 1.44-1.55; $P < .001$).

Statistically significant between-sample inconsistency was found, with an $I^2$ value of 75.7% (Cochran Q = 96.57, $P < .001$). The Egger test ($z = 4.27$, $P < .001$) suggested that publication bias was
statistically significant. Trim-and-fill analysis estimated 3 missing studies on the left side of eFigure 1 in the Supplement, indicating that negative studies may have not been reported. However, our results remained statistically significant when the putative missing studies were included (SMD = 1.37 [95% CI, 1.07-1.67]; \( P < .001 \)). Meta-regressions of effect sizes against age (n = 24) \( 5,29,33,36,38,40-43,48-53,55-57,59,61 \) and sex (n = 22) \( 25,29,31,33,36,38,40-43,48,50-53,55-57,59,61 \) showed that neither factor was a statistically significant moderator of effect sizes.

Ketamine Preparation
Both racemic ketamine and s-ketamine preparations resulted in a statistically significant increase in total symptom scores compared with placebo. Large effect sizes were found for racemic ketamine (SMD = 1.40 [95% CI, 1.12-1.68]; \( P < .001 \)) and s-ketamine (SMD, 2.03 [95% CI, 1.15-2.92]; \( P < .001 \)). There was no significant difference between the methods on the magnitude of the effect size.

Blinding Method
Unblinded or single-blind methods (SMD = 1.71 [95% CI, 1.02-2.39]; \( P < .001 \)) and double-blind methods (SMD = 1.45 [95% CI, 1.15-1.75]; \( P < .001 \)) both resulted in a statistically significant association of the ketamine condition with total symptoms. There was no significant difference between the methods on the magnitude of the effect size.

Infusion Method
Bolus and a continuous infusion (SMD = 1.55 [95% CI, 1.23-1.88]; \( P < .001 \)) and a continuous infusion only (SMD = 1.27 [95% CI, 0.73-1.81]; \( P < .001 \)) were both associated with a statistically significant increase in total symptoms. There was no significant difference between the methods on the magnitude of the effect size.

Single-Day vs Multiple-Day Studies
Two single-day studies \( ^{29,30} \) (SMD = 2.29 [95% CI, 1.69-2.89]; \( P < .001 \)) and 17 multiple-day studies \( ^{25,31,36,38,40-43,48,50-53,55,56,61} \) (SMD = 1.39 [95% CI, 1.12-1.66]; \( P < .001 \)) were associated with a statistically significant increase in total symptoms. Studies in which ketamine and placebo conditions were conducted on the same day found a significantly greater magnitude of effect (effect size, 2.29 [95% CI, 1.69-2.89] vs 1.39 [95% CI, 1.12-1.66]; \( P = .007 \)) (eFigure 4 in the Supplement).

Positive Psychotic Symptoms
Positive symptom scores were analyzed using data from 21 studies \( ^{3,28,35,39,41,44-47,49,51,58-60} \) consisting of 513 healthy participants exposed to the ketamine and placebo conditions. Positive symptom scores were transiently increased in the ketamine condition compared with the placebo condition (SMD = 1.55 [95% CI, 1.29-1.81]; \( P < .001 \)) (Figure 3). The result remained statistically significant in all iterations of the leave-one-out analysis (SMD range, 1.47-1.60; \( P < .001 \)). Statistically significant between-sample inconsistency was found, with an \( I^2 \) value of 74.9% (Cochran Q = 81.40; \( P < .001 \)). Findings of the Egger test (\( z = 5.06; P < .001 \)) suggested that publication bias was significant. Trim-and-fill analysis estimated 1 missing study on the left side (eFigure 2 in the Supplement). Results remained statistically significant with putative missing studies included (SMD = 1.49 [95% CI, 1.18-1.80]; \( P < .001 \)). Meta-regressions of effect sizes against age (n = 20) \( 3,29,35,39,41,44-47,49,51,58-60 \) or sex (n = 19) \( 3,28,31,33,35,39,41,44-47,50,51,58-60 \) showed that neither was a statistically significant moderator of effect sizes.

Ketamine Preparation
Both racemic ketamine (SMD = 1.50 [95% CI, 1.17-1.82]; \( P < .001 \)) and s-ketamine (SMD = 1.70 [95% CI, 1.23-2.18]; \( P < .001 \)) preparations resulted in a statistically significant increase in positive symptom scores compared with placebo, both with large effect sizes. There was no significant difference between the methods on the magnitude of the effect size.
Blinding Method
Unblinded or single-blind (SMD = 1.32 [95% CI, 0.96-1.67]; \(P < .001\)) and double-blind (SMD = 1.68 [95% CI, 1.30-2.07]; \(P < .001\)) methods resulted in a statistically significant effect of ketamine condition on the positive symptoms. However, there was no significant difference in the magnitude of the effect between the 2 methods.

Infusion Method
Both a bolus followed by continuous infusion method (n = 19)\(^{28,35,39,41,44,45,49-51,58-60,64}\) (SMD = 1.63 [95% CI, 1.36-1.90]; \(P < .001\)) and a continuous infusion alone (n = 2)\(^{46,47}\) (SMD = 0.84 [95% CI, 0.35-1.33]; \(P < .008\)) induced a statistically significant increase in positive symptoms. However, studies using a bolus and continuous infusion method induced a statistically significantly greater magnitude of effect (effect size, 1.63 [95% CI, 1.36-1.90]) compared to continuous infusion alone (effect size, 0.84 [95% CI, 0.35-1.33]; \(P = .006\)) (eFigure 5 in the Supplement).

Single-Day vs Multiple-Day Studies
Single-day (SMD = 1.54 [95% CI, 1.19-1.89]; \(P < .001\)) and multiple-day (SMD = 1.53 [95% CI, 1.15-1.90]; \(P < .001\)) studies both resulted in a statistically significant increase in positive symptoms. There was no significant difference in the magnitude of the effect between the 2 methods.

Negative Symptoms
Negative symptom scores were analyzed using data from 22 studies\(^{3,25,28-30,32-36,41,44-47,49-51,54,58-60}\) consisting of 527 healthy participants exposed to the ketamine and placebo conditions. Negative symptom scores were transiently increased in the ketamine condition compared with the placebo condition (SMD = 1.16 [95% CI, 0.96-1.35]; \(P < .001\)) (Figure 4). The result remained statistically significant in all iterations of the leave-one-out analysis (SMD range, 1.11-1.19; \(P < .001\)).

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**Figure 3. Standardized Mean Difference (SMD) in Positive Symptom Scores for Healthy Volunteers After Ketamine vs Placebo Administration**

| Source                  | SMD (95% CI) | Increased symptoms in placebo condition | Increased symptoms in ketamine condition |
|-------------------------|-------------|------------------------------------------|-----------------------------------------|
| Breier et al,\(^{35}\) 1997 | 1.61 (0.89 to 2.33) |                                            |                                          |
| Krystal et al,\(^{39}\) 1998    | 2.12 (1.39 to 2.86) |                                            |                                          |
| Newcomer et al,\(^{31}\) 1999  | 1.18 (0.52 to 1.84) |                                            |                                          |
| Krystal et al,\(^{46}\) 1999    | 4.66 (3.15 to 6.17) |                                            |                                          |
| Anand et al,\(^{44}\) 2000    | 1.76 (0.98 to 2.55)  |                                            |                                          |
| Micallef et al,\(^{47}\) 2002  | 0.50 (-0.24 to 1.23) |                                            |                                          |
| Krystal et al,\(^{46}\) 2003    | 1.02 (0.55 to 1.50)  |                                            |                                          |
| Krystal et al,\(^{35}\) 2005    | 1.25 (0.74 to 1.75)  |                                            |                                          |
| Krystal et al,\(^{36}\) 2006    | 1.06 (0.62 to 1.50)  |                                            |                                          |
| Dickerson et al,\(^{41}\) 2010 | 0.79 (0.56 to 1.02)  |                                            |                                          |
| Nagel et al,\(^{37}\) 2011    | 2.25 (1.30 to 3.21)  |                                            |                                          |
| D’Souza et al,\(^{34}\) 2012   | 1.97 (1.38 to 2.57)  |                                            |                                          |
| Vernaleken et al,\(^{36}\) 2013 | 1.04 (0.27 to 1.81)  |                                            |                                          |
| Driesen et al,\(^{60}\) 2013    | 2.07 (1.33 to 2.81)  |                                            |                                          |
| Powers et al,\(^{38}\) 2015    | 1.32 (0.71 to 1.94)  |                                            |                                          |
| Höflich et al,\(^{47}\) 2015   | 1.58 (1.04 to 2.11)  |                                            |                                          |
| Kleinloog et al,\(^{28}\) 2015 | 2.67 (1.91 to 3.44)  |                                            |                                          |
| Kraguljac et al,\(^{40}\) 2017 | 1.61 (0.84 to 2.37)  |                                            |                                          |
| Theibes et al,\(^{29}\) 2017   | 1.24 (0.71 to 1.77)  |                                            |                                          |
| Gent-`t Jong et al,\(^{33}\) 2018 | 1.70 (0.88 to 2.52)  |                                            |                                          |
| D’Souza et al,\(^{34}\) 2018   | 1.67 (1.08 to 2.27)  |                                            |                                          |
| Random-effects model for all studies | 1.55 (1.29 to 1.81) |                                            |                                          |

Scores include Positive and Negative Syndrome Scale and Brief Psychiatric Rating Scale. A statistically significant increase in positive symptoms occurred in healthy volunteers in the ketamine condition compared with the placebo condition (SMD = 1.55 [95% CI, 1.29-1.81]; \(P < .001\)). Each square shows the effect size for a single study, with the horizontal error bars representing the width of the 95% CI. The size of the squares reflects the weight attributed to each study. The diamond illustrates the summary effect size, and the width of the diamond depicts the width of the overall 95% CI.
Statistically significant between-sample inconsistency was found, with an $I^2$ value of 64.6% (Cochran Q = 66.55; P < .001). Findings of the Egger test ($z = 5.12; P < .001$) suggested that publication bias was significant. Trim-fill analysis estimated 2 missing studies on the left side of eFigure 3 in the Supplement. Results remained statistically significant with putative missing studies included (SMD = 1.09 [95% CI, 0.89-1.30]; P < .001). Meta-regressions of effect sizes against age ($n = 20$) $3,25,30,32-36,44-47,49-51,58-60$ or sex ($n = 19$) $3,25,28-30,33-36,44-47,50,51,58-60$ showed that neither was a statistically significant moderator of effect sizes.

Ketamine Preparation

Both racemic ketamine (SMD = 1.13 [95% CI, 0.90-1.36]; P < .001) and s-ketamine (SMD = 1.25 [95% CI, 0.86-1.64]; P < .001) preparations resulted in a statistically significant transient increase in negative symptom scores compared with placebo, with large effect sizes. There was no significant difference between the methods on the magnitude of the effect size.

Blinding Method

Unblinded or single-blind (SMD = 0.98 [95% CI, 0.63-1.34]; P < .001) and double-blind (SMD = 1.29 [95% CI, 1.09-1.50]; P < .001) methods resulted in a statistically significant association of the ketamine condition with negative symptoms. There was no significant difference between the methods on the magnitude of the effect size.

Infusion Method

Both bolus and a continuous infusion (SMD = 1.19 [95% CI, 0.96-1.41]; P < .001) and a continuous infusion only (SMD = 1.06 [95% CI, 0.71-1.40]; P < .001) were associated with a statistically significant increase in negative symptoms. There was no significant difference between the methods on the magnitude of the effect size.

**Figure 4. Standardized Mean Difference (SMD) in Negative Symptom Scores in Healthy Volunteers After Ketamine vs Placebo Administration**

| Source | SMD (95% CI) | Increased symptoms in placebo condition | Increased symptoms in ketamine condition |
|--------|--------------|----------------------------------------|-----------------------------------------|
| Malhotra et al,25 1997 | 1.35 (0.67 to 2.02) | | |
| Krystal et al,35 1998 | 1.51 (0.91 to 2.11) | | |
| Krystal et al,45 1999 | 2.21 (1.40 to 3.02) | | |
| Anand et al,44 2000 | 1.38 (0.69 to 2.06) | | |
| Duncan et al,46 2001 | 0.99 (0.39 to 1.59) | | |
| Micalefe et al,47 2002 | 0.89 (0.07 to 1.71) | | |
| Krystal et al,48 2003 | 1.16 (0.67 to 1.66) | | |
| Krystal et al,5 2005 | 1.54 (0.99 to 2.10) | | |
| Krystal et al,16 2006 | 1.11 (0.66 to 1.56) | | |
| Dickerson et al,41 2010 | 0.51 (0.29 to 0.73) | | |
| Nagel et al,52 2011 | 2.30 (1.33 to 3.26) | | |
| D’Souza et al,53 2015 | 1.48 (0.98 to 1.98) | | |
| Vernallengen et al,54 2013 | 0.86 (0.13 to 1.58) | | |
| Driesen et al,55 2013 | 0.67 (0.21 to 1.13) | | |
| Powers et al,56 2015 | 1.09 (0.52 to 1.65) | | |
| Höflich et al,57 2015 | 0.73 (0.32 to 1.13) | | |
| Kleinloog et al,58 2015 | 1.23 (0.75 to 1.70) | | |
| Krakujac et al,59 2017 | 1.95 (1.09 to 2.82) | | |
| Thiebes et al,60 2017 | 1.40 (0.84 to 1.97) | | |
| Grent ‘t Jong et al,61 2018 | 1.59 (0.80 to 2.38) | | |
| Abdallah et al,62 2018 | 0.45 (-0.10 to 1.00) | | |
| D’Souza et al,63 2018 | 0.76 (0.32 to 1.19) | | |
| Random-effects model for all studies | 1.16 (0.96 to 1.35) | | |

Scores include Positive and Negative Syndrome Scale and Brief Psychiatric Rating Scale. A statistically significant increase in negative symptoms occurred in healthy volunteers in the ketamine condition compared with the placebo condition (SMD = 1.16 [95% CI, 0.96-1.35]; P < .001). Each square shows the effect size for a single study, with the horizontal error bars representing the width of the 95% CI. The size of the squares reflects the weight attributed to each study. The diamond illustrates the summary effect size, and the width of the diamond depicts the width of the overall 95% CI.
Single-Day vs Multiple-Day Studies

Single-day (SMD = 1.01 [95% CI, 0.56-1.47]; \( P < .001 \)) and multiple-day (SMD = 1.16 [95% CI, 0.94-1.39]; \( P < .001 \)) studies both resulted in a statistically significant increase in negative symptoms. However, there was no significant difference in the magnitude of the effect between the 2 methods.

Comparison of Positive and Negative Effect Sizes

A comparison of effect sizes demonstrated that the ketamine condition had a greater association with positive symptoms compared with negative symptoms (estimate, 0.36 [95% CI, 0.12-0.61]; \( z = 2.90; P = .004 \)).

Subanalyses of BPRS and PANSS scales are presented in eMethods 10 in the Supplement. In summary, there was no significant difference between the 2 measures for any of the symptom domains (total, positive, and negative). Inconsistency analyses for the subanalyses are presented in eMethods 9 in the Supplement.

Effects of Ketamine in Patients With Schizophrenia

After 7 studies with overlapping data sets were excluded, \( ^{64-70} \) 3 studies were included in the analysis of patients with schizophrenia. \( ^{25-27} \) No meta-analysis was possible because there were an insufficient number of papers. Studies with change scores were included in this section of the review.

Two studies \( ^{25,26} \) examined the association of acute ketamine administration on total BPRS scores in patients with schizophrenia, and both found that ketamine was associated with a statistically significant increase in total BPRS scores. Two studies \( ^{26,27} \) investigated the association of ketamine administration with positive and negative BPRS scores in patients with schizophrenia. Both found ketamine was associated with a statistically significant transient increase in positive symptoms. One study \( ^{27} \) found ketamine was associated with a statistically significant increase in negative symptoms, but the other study \( ^{26} \) to assess this factor did not find a significant association of ketamine with negative symptoms. The findings of these studies are summarized in eTable 4 in the Supplement.

Risk of Bias Across Studies

Eight studies \( ^{29,41,50,51,54,55,57,58} \) had a high risk of bias when the Newcastle-Ottawa tool was used to assess bias, mainly owing to not documenting certain aspects of the design protocol and therefore losing a point for being unclear. The Cochrane tool for assessment of bias across studies highlighted an unclear risk of bias across the selection bias domain but low risk across all other domains (performance bias, detection bias, attrition bias, and reporting bias) (eMethods 2-4 in the Supplement).

Discussion

Our main findings were that acute ketamine administration was associated with a large effect size for increases in positive, negative, and total symptom scores in healthy volunteers. Moreover, ketamine was associated with greater increases in positive symptoms than in negative symptoms.

Insufficient studies were available to conduct a meta-analysis of the association of ketamine with psychopathology in schizophrenia. Although transient increases in positive, negative, and total symptoms in patients with schizophrenia were reported, given the limited data, firm conclusions on effects in schizophrenia cannot be drawn, and further studies are needed. These findings extend the understanding of the symptoms associated with ketamine by showing that either racemic ketamine or s-ketamine are associated with positive, negative, and total symptoms in healthy volunteers with very large effect sizes across study settings and designs. To give some clinical context to the increased effect sizes seen with ketamine administration, the average mean difference in the total PANSS scores between the placebo and ketamine conditions was 18.40. Were this increase in symptom rating to occur in a patient with schizophrenia, it would approximately equate to a change...
from mild illness severity to markedly ill on the Clinical Global Impression Scale and represent a clinically meaningful increase in symptoms.71

This study identifies high levels of between-study inconsistency. Our subgroup analyses indicate that this inconsistency could be owing to study design factors. Specifically, studies that used a bolus plus infusion protocol showed larger increases (approximately double) in positive symptoms than those using only a continuous infusion. Moreover, studies that administered ketamine and placebo on the same day found a greater increase in total symptoms. The first finding could be owing to a faster time to and/or higher peak concentration of ketamine, consistent with a study showing a positive association between ketamine concentration and symptom induction.28 It is less clear why giving ketamine and placebo on the same day was associated with greater induction of symptoms, but this factor could reflect unblinding because one study was unblinded29 and the other was single blinded with the condition order randomized.30 Another explanation might be that both conditions on the same day controls better for the day-to-day variance that may occur in mood and biology. When heterogeneity was assessed for each individual subgroup, it was moderate to high for most analyses, suggesting that these subgroups did not account for all of the inconsistency seen within the meta-analysis.

Association of Age and Sex With Ketamine-Induced Psychopathology
Neither age nor sex were associated with the severity of psychotic symptoms induced by ketamine in healthy volunteers. However, the studies included in our meta-analysis only include adults (range of mean ages, 22-40 years). In children, fewer ketamine-induced symptoms might occur because children are less likely to experience psychotic symptoms than adults when given ketamine for anesthesia.1 However, animal studies find that ketamine has a greater neurotoxic effect in the period from puberty to early adulthood.72 Sex did not moderate the magnitude of effect for any of the symptom measures in our study, consistent with findings by Morgan and colleagues73 but in contrast with preclinical evidence that female rats are more susceptible than male rats to the neurotoxic74 and behavioral75 effects of ketamine. This difference between clinical and preclinical evidence may reflect the higher doses used in the animal studies (5-180 mg/kg) compared with humans (approximately 0.65 mg/kg).

Implications for Future Study Design and Reporting
Our findings are of particular relevance for the therapeutic use of ketamine and for future study design. First, we provide evidence that the use of bolus plus continuous infusion is associated with larger transient psychotomimetic effects. Second, inadequate reporting of methods precluded our ability to test the effects of other key methodological factors. One recommendation from our findings is therefore for future studies to report methods with greater detail to enable these factors to be investigated and aid replication.76 Details of specific relevance to studies such as these include the dose of ketamine and fasting status before receiving ketamine.

Ketamine Model of Schizophrenia
We found that ketamine was associated with the induction of both transient positive and negative symptoms of schizophrenia in healthy people and with worsened symptoms in patients with schizophrenia. To the extent that any drug can model a complex disorder such as schizophrenia, the results of this meta-analysis support the use of ketamine to model schizophrenia-like or psychotomimetic symptoms and suggest that it provides a more comprehensive model of schizophrenia than drugs such as amphetamine, which does not reliably induce negative symptoms.3 However, we found that the induction of negative symptoms is statistically significantly less marked than that of positive psychotic symptoms in healthy people, and it was only seen in 1 of the 2 studies in schizophrenia.27 The negative symptom analysis had an extra study with a continuous infusion method. This difference may have reduced the effect because the continuous infusion method appears less likely than the bolus and continuous method to induce psychotic symptoms. However,
the psychotic symptom analysis had more unblinded studies and fewer studies that completed both conditions on the same day. Notwithstanding these methodological considerations, this finding suggests that acute ketamine administration is associated with more positive than negative symptoms, although the magnitude of negative symptoms associated with ketamine is still large.

**Implications for Therapeutic Use of Ketamine**

Ketamine is being evaluated as a treatment for depression and some other disorders. Our findings highlight the potential risk that ketamine may induce transient positive (psychotic), negative, and other symptoms, particularly because the dose and route used to treat depression (approximately 0.5 mg/kg intravenously) is similar to those used in studies in this meta-analysis. Evidence suggests that ketamine can induce perceptual disturbances and psychotic symptoms in patients with depression (mean BPRS score, 12.6) with slightly higher positive BPRS scores than those seen in the studies included in this meta-analysis (average mean score across all BPRS studies, 7.5). People with a history of psychosis may be more vulnerable to the effects of ketamine. Our finding that using a bolus and continuous infusion method increases the effect of ketamine on psychotic symptoms highlights the importance of using slower infusions (40-60 minutes) of ketamine, an approach now adopted by some therapeutic trials.

**Strengths and Limitations**

Strengths of this study include the relatively large sample size and inclusion of additional data provided by authors. However, there was significant inconsistency in the summary effect sizes, suggesting variability in effects between studies. This factor can be explained in part by differences in study design, as indicated by our sensitivity findings described above. We cannot explore the effect of other methodological differences that may contribute to inconsistency, such as differences in ketamine doses or fasting status, because few studies reported sufficient detail to allow this subanalysis. The inconsistency in total symptom score may also be explained by the inclusion of different BPRS versions. A mixture of 16, 18, 20, and 24 total item scales were used, and very few studies made it clear which items they included. Consequently, individual subgroup analyses of different versions could not be conducted.

There were several important differences in exclusion criteria between the studies. In particular, most of the studies did not exclude concurrent use of psychotropic drugs and only a few studies excluded participants with prior ketamine exposure. Although some evidence suggests that repeated ketamine exposure does not cause behavioral sensitization in humans, it would have been useful to have examined these data in more depth to determine whether these factors may alter results owing to drug tolerance or differences in subjective experience due to familiarity with prior exposure. Nevertheless, we used a random-effects model, which is a robust method of calculating the effect size when there is statistically significant inconsistency between studies.

Interestingly, the blinding status did not alter the magnitude of the effect size for total, positive, or negative symptoms. Blinding participants in these experiments may be very difficult because the dissociative anesthetic effects of ketamine can be very obvious to both participant and study personnel. This possibility may further explain the heterogeneity of results because the participants’ expectations may have contributed to their drug response. Future studies could include a low dose of ketamine or active comparator, such as midazolam hydrochloride, to address this important question.

We aimed to determine ketamine’s maximal ability to induce psychotomimetic symptoms. Where symptom scales were reported at different time points, we selected the point with the highest ketamine-induced symptom score. Where this occurred, the symptom measure for the placebo group was taken at the corresponding point. Where studies included different concentrations of ketamine, we used the highest dose, again using the symptom score for the corresponding placebo condition. Therefore, the effect sizes in this study are likely to be the largest effect size seen with
ketamine. Further work is thus required to better characterize the dose-response relationship and
time course of ketamine’s psychotomimetic affects.

Conclusions

We provide meta-analytic evidence that ketamine is associated with the induction of transient
positive, negative, and total symptoms, with a greater increase in positive than negative symptoms
in healthy volunteers. These findings support the use of ketamine as a pharmacological model of
schizophrenia and, given that using a bolus plus continuous infusion method leads to greater positive
psychotic symptoms, indicate that the bolus plus infusion is the best approach for this model.
Ketamine is used to treat pain and for major depression. Our findings indicate a potential risk of
ketamine inducing schizophreniform symptoms when it is used for these indications and that a slow
infusion without bolus is preferable to minimize these risks. Further research is needed to determine
the risk of these effects when ketamine is used therapeutically.
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SUPPLEMENT.
eMethods 1. Studies and Data Not Included in Meta-analysis

eMethods 2. Newcastle-Ottawa Assessment Scale for Cohort Studies

eMethods 3. Cochrane Tool for Assessment of Bias Within Individual Studies

eMethods 4. Cochrane Tool for Assessment of Bias Across Studies

eMethods 5. Refitting the Model Using r’s Taking Values 0.1

eMethods 6. Refitting the Model Using r’s Taking Values 0.7

eMethods 7. Comparison of the Effect of Ketamine on Positive and Negative Symptoms Using Correlation Coefficient of 0.1

eMethods 8. Comparison of the Effect of Ketamine on Positive and Negative Symptoms Using Correlation Coefficient of 0.7

eMethods 9. Heterogeneity Statistics for Subgroup Analyses

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eFigure 1. Funnel Plot for Total Symptoms
eFigure 2. Funnel Plot for Positive Symptoms

eFigure 3. Funnel Plot for Negative Symptoms

eFigure 4. Subgroup Analysis of Single-Day vs Multiple-Day Studies for Total Symptoms

eFigure 5. Subgroup Analysis of Method of Infusion (Bolus and a Continuous Infusion vs Only a Continuous Infusion) Positive Symptoms

eTable 1. Raw Data Used in Total BPRS and PANSS Analysis for Healthy Participants

eTable 2. Raw Data Used in Positive BPRS and PANSS Analysis for Healthy Participants

eTable 3. Raw Data Used in Negative BPRS and PANSS Analysis for Healthy Participants

eTable 4. Raw Data used in Total, Negative and Positive BPRS and PANSS Analysis for People With Schizophrenia

eTable 5. Study Description, Ketamine Method, Placebo Condition, Symptoms (BPRS and PANSS), and Exclusion Criteria in Studies Examining Acute Ketamine Administration in Healthy Controls

eTable 6. Study Description, Ketamine Administration Method, Placebo Condition, Symptoms (BPRS), and Exclusion Criteria in Studies Examining Acute Ketamine Administration to Patients With Schizophrenia

eReferences.