Effect of mechanical ventilation in the prone position on clinical outcomes in patients with acute hypoxemic respiratory failure: a systematic review and meta-analysis

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Abstract

Background: Mechanical ventilation in the prone position is used to improve oxygenation in patients with acute hypoxemic respiratory failure. We sought to determine the effect of mechanical ventilation in the prone position on mortality, oxygenation, duration of ventilation and adverse events in patients with acute hypoxemic respiratory failure.

Methods: In this systematic review we searched MEDLINE, EMBASE, the Cochrane Central Register of Controlled Trials and Science Citation Index Expanded for articles published from database inception to February 2008. We also conducted extensive manual searches and contacted experts. We extracted physiologic data and clinically relevant outcomes.

Results: Thirteen trials that enrolled a total of 1559 patients met our inclusion criteria. Overall methodologic quality was good. In 10 of the trials (n = 1486) reporting this outcome, we found that prone positioning did not reduce mortality among hypoxemic patients (risk ratio [RR] 0.96, 95% confidence interval [CI] 0.84–1.09; p = 0.52). The lack of effect of ventilation in the prone position on mortality was similar in trials of prolonged prone positioning and in patients with acute lung injury. In 8 of the trials (n = 633), the ratio of partial pressure of oxygen to inspired fraction of oxygen on day 1 was 34% higher among patients in the prone position than among those who remained supine (p < 0.001); these results were similar in 4 trials on day 2 and in 5 trials on day 3. In 9 trials (n = 1206), the ratio in patients assigned to the prone group remained 6% higher the morning after they returned to the supine position compared with patients assigned to the supine group (p = 0.07). Results were quantitatively similar but statistically significant in 7 trials on day 2 and in 6 trials on day 3 (p = 0.001). In 5 trials (n = 1004), prone positioning was associated with a reduced risk of ventilator-associated pneumonia (RR 0.81, 95% CI 0.66–0.99; p = 0.04) but not with a reduced duration of ventilation. In 6 trials (n = 504), prone positioning was associated with an increased risk of pressure ulcers (RR 1.36, 95% CI 1.07–1.71; p = 0.01). Most analyses found no to moderate between-trial heterogeneity.

Interpretation: Mechanical ventilation in the prone position does not reduce mortality or duration of ventilation despite improved oxygenation and a decreased risk of pneumonia. Therefore, it should not be used routinely for acute hypoxemic respiratory failure. However, a sustained improvement in oxygenation may support the use of prone positioning in patients with very severe hypoxemia, who have not been well-studied to date.

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and new atelectasis in the now dependent ventral regions is comparatively minor. In addition, lung perfusion in the prone position is more homogeneous. Shunt conditions are therefore reduced and ventilation is better matched to perfusion. Other clinical effects of prone positioning may include enhanced postural drainage of secretions,\textsuperscript{8,10} decreasing the risk of ventilator-associated pneumonia. Effects may also include decreased alveolar overdistension, cyclic alveolar collapse and ventilator-induced lung injury.\textsuperscript{11} For this reason, some investigators have recommended prone positioning for mechanical ventilation in the treatment of acute lung injury.\textsuperscript{8,11}

Although ventilation in the prone position offers physiologic advantages and does not require specialized tools, one survey found that in most ICUs, 3 personnel (range 2–6) were required to turn an adult patient.\textsuperscript{13} These caregivers must handle major safety challenges in putting patients with life-threatening hypoxemia in the prone position, including disconnection or removal of endotracheal tubes or intravascular catheters, and kinking or secretion-induced plugging of endotracheal tubes.\textsuperscript{13}

Despite prone positioning’s physiologic advantages, individual randomized controlled trials have not demonstrated its superior clinical outcomes compared with supine positioning. Consequently, we conducted a systematic review and meta-analysis to evaluate the effect of prone positioning on clinical outcomes, including mortality, oxygenation, ventilator-associated pneumonia, duration of ventilation and adverse events, in patients with acute hypoxic respiratory failure.

**Methods**

**Literature search**

We searched MEDLINE, EMBASE, the Cochrane Central Register of Controlled Trials and Science Citation Index Expanded for articles published from database inception to February 2008. Our search strategy is outlined in Appendix 1, available at www.cma.ca/cgi/content/full/178/9/1153/DC2. We also searched supplementary data sources, including the “related articles” feature on PubMed; bibliographies of included studies and review articles; conference proceedings of the American Thoracic Society (1994–2007), the American College of Chest Physicians (1994–2007), the European Society of Intensive Care Medicine (1994–2007) and the Society of Critical Care Medicine (1994–2008); and clinical trial registries (www.clinicaltrials.gov, www.controlled-trials.com). We contacted clinical experts and the authors of all included studies for additional data. We did not impose language restrictions.

**Study selection**

We included studies that met 3 criteria. First, they enrolled adult or pediatric patients with acute hypoxic respiratory failure (defined as the ratio of partial pressure of oxygen to inspired fraction of oxygen ≤ 300 mm Hg), including acute lung injury and acute respiratory distress syndrome, who received mechanical ventilation. Second, they randomly assigned patients to 2 or more groups, including a treatment group that received ventilation at least once in the prone position and a control group that received ventilation in the supine position. Third, they reported all-cause mortality, the ratio of partial pressure of oxygen to inspired fraction of oxygen, ventilator-associated pneumonia, the duration of ventilation, the number of ventilator-free days from randomization to day 28 or 30, or adverse events, including pressure ulcers, endotracheal tube obstruction, unplanned extubation, dislodgement of central venous catheters or thoracostomy tubes, pneumothoraces and cardiac arrests. Our quantitative analyses included trials that enrolled adults or postneonatal children. Excluding the pediatric trials did not change any results; therefore, we present only the combined results.

We also considered trials that assigned patients in alternating fashion or by hospital registry number (quasi-randomization), or involved cointerventions, such as high-frequency oscillation or nitric oxide, that were specified as part of the intervention and were applied equally to both groups. We used authors’ definitions of acute lung injury and acute respiratory distress syndrome. We excluded randomized crossover trials that assigned patients to both treatment and control groups.

**Data abstraction and validity assessment**

Each of us independently evaluated studies for inclusion and abstracted data on study methods and outcomes; disagreements were resolved by consensus. Measures of study quality included method and concealment of allocation (adequate v. inadequate), postallocation withdrawals (yes v. no), patients with missing mortality status owing to loss to follow-up (yes v. no), crossovers between groups (yes v. no), analysis of data by group to which patients were originally assigned (yes v. no), blinding of outcome assessors for ventilator-associated pneumonia (blinded or centrally adjudicating assessors v. neither), cointerventions (standardization or equal application of mechanical ventilation, ventilator weaning, sedation and paralysis, and alternative treatments for hypoxemia), and early stopping of the trial before planned enrolment was completed (yes v. no).

We contacted authors of all included trials to clarify methodology and request data missing from prespecified analyses.

**Statistical analysis**

The primary outcome was all-cause mortality in the ICU at any time after randomization; if ICU mortality was not reported, we used mortality at 28 or 30 days after randomization or hospital mortality. A priori, we planned subgroup analyses based on patient population (acute lung injury or acute respiratory distress syndrome v. other) and duration of prone positioning (prolonged, which we defined as up to 24 hours daily for more than 2 days, v. short-term).

Secondary outcomes included ventilator-associated pneumonia, the number of days on mechanical ventilation and ventilator-free days, oxygenation on days 1–3 and adverse events. Oxygenation outcomes are presented only for days 1–3 because the extent of missing data for subsequent days (in trials reporting these outcomes) limits the interpretability of these analyses. To show the maximal effect of prone positioning on oxygenation, we compared the mean ratio of partial pressure of oxygen to inspired fraction of oxygen, measured in the prone group at the end of a prone manoeuvre,
with the simultaneously recorded measurement in the supine group for each day. For this measurement, day 1 refers to the end of the first proning session (for sessions lasting less than 24 hours) or the end of 24 hours (for continuous proning lasting longer than 24 hours). To measure the difference in oxygenation that remained after patients in the prone group were returned to the supine position, we compared the mean ratios of partial pressure of oxygen to inspired fraction of oxygen between the prone and supine groups that were measured in the morning (just before the subsequent proning manoeuvre in the prone group). Day 1 for this measurement refers to the measurement taken the morning after the first proning session and applies only to proning sessions that lasted less than 24 hours. Finally, in a post hoc analysis we compared the mean ratio of partial pressure of oxygen to inspired fraction of oxygen measured within 1 hour of patients being turned to the prone position with the near-simultaneous measurement in the supine group.

In our meta-analysis, all statistical tests were 2-sided, and we considered \( p < 0.05 \) to be statistically significant. We report continuous outcomes as weighted mean differences (a measure of absolute change) for number of days of mechanical ventilation and ventilator-free days, and as ratios of means (a measure of relative change)\(^{14,15}\) for ratio of partial pressure of oxygen to inspired fraction of oxygen. We report binary outcomes (mortality, ventilator-associated pneumonia and adverse events) as risk ratios (RRs). All outcomes are presented with 95% confidence intervals (CIs). At least 2 of us independently conducted each analysis to minimize data management errors.\(^{25}\)

We measured heterogeneity and expressed it as \( I^2 \), the percentage of total variation across studies owing to between-study heterogeneity rather than chance.\(^{16,17}\) with suggested thresholds for low (\( I^2 = 25\%–49\%) \), moderate (\( I^2 = 50\%–74\%) \) and high (\( I^2 \geq 75\%) \) values.\(^{17}\) A priori hypotheses to explain moderate to high heterogeneity in mortality and ventilator-associated pneumonia included study population and duration of prone positioning, and study quality (adequate v. inadequate concealment of patient assignment, including quasi-randomization, and blinded or centrally adjudicating assessors v. neither [for ventilator-associated pneumonia only]). We performed meta-analyses using a random-effects model, which incorporates within- and between-study variation and provides more conservative treatment estimates when heterogeneity is present.\(^{19}\)

To assess publication bias, we visually examined the funnel plot for mortality (standard error of the natural logarithm of RR v. RR for each trial) and performed a Begg rank correlation test\(^{29}\) and Macaskill regression test,\(^{30}\) as modified by Peters and colleagues,\(^{21}\) in which we considered \( p < 0.10 \) to be statistically significant.

**Results**

**Literature search**

Through the searches of bibliographic databases we identified 1676 citations. Of these citations and additional citations retrieved from other sources, we retrieved 50 studies for detailed evaluation and excluded 29. We selected 13 primary randomized and quasi-randomized trials for inclusion in our review and meta-analysis,\(^{22–34}\) along with 8 references providing duplicate or additional data. Of the studies retained for analysis, 12 trials contributed oxygenation data, 10 contributed data on adverse events, and 10 were included in our primary mortality analysis.

**Appendix 2 (available at www.cmaj.ca/cgi/content/full/178/9/1153/DC2) describes the 13 included trials,\(^{22–34}\) which enrolled a total of 1559 patients (median per trial 28, range 16–802) with acute lung injury,\(^{73,26,28,30}\) acute respiratory distress syndrome,\(^{22,29,31,32,34}\) or acute hypoxemic respiratory failure.\(^{25,27,33}\) The largest trial (\( n = 802 \)) enrolled patients with acute hypoxemic respiratory failure; 51% had acute lung injury or acute respiratory distress syndrome and 7% had cardiogenic pulmonary edema.\(^{27}\) Most of the trials enrolled patients within

![Figure 1: Flow of studies in the systematic review. *The records were identified in electronic database search. ‡The records retrieved for more detailed evaluation came from the electronic databases and other sources. ‡‡We included 13 primary trials and 8 references with duplicate or additional data. Of the studies retained for analysis, 12 trials contributed oxygenation data, 10 contributed data on adverse events, and 10 were included in our primary mortality analysis.](https://www.cmaj.ca/content/full/178/9/1153/DC2)
48 hours of diagnosis. One trial enrolled comatose patients (Glasgow coma score ≤ 9), who were not necessarily in acute hypoxemic respiratory failure, within 24 hours of mechanical ventilation. Only data for subgroups of patients with hypoxemia (ratio of partial pressure of oxygen to inspired fraction of oxygen ≤ 300 mm Hg), or acute lung injury or acute respiratory distress syndrome were used in our analyses. Two trials enrolled only children. 28,33

Patients received mechanical ventilation in the prone position for a median of 12 hours per day (range 4–24 hours), and proning manoeuvres continued either for a prespecified period 22,25,26,29,31 or until prespecified clinical improvements occurred 23,24,27,28,30,31 (median duration of proning 4 days, range 1–10 days). About 2–6 clinical personnel 23,25,28,31 were required for each turning procedure (4–6 personnel were required in trials enrolling adults). 23,25,31 The turning procedure lasted a mean of 10 (standard deviation 12) minutes in the only trial that reported this information. 29

All but 1 trial, 22 which was available only in abstract form, provided some description of mechanical ventilation. Five trials mandated low tidal volume ventilation (6–8 mL/kg body weight), 20,28,30,32,34 4 trials used protocols to adjust positive end-expiratory pressure, 20,28,32,34 and 5 trials reported a mean positive end-expiratory pressure of 7–12 cm H2O during the study period. 23,27,28,30,31 Additional cointerventions that were specified as part of the treatment and applied to all patients in both groups included use of nitric oxide 33 and high-frequency oscillation. 29

The trials had high methodologic quality (Appendix 2). Most described adequate allocation concealment (9 trials), 22–24,27–29 Allocation was not concealed in 3 trials, 25,33,34 of which 2 assigned patients using alternate allocation. 25,33 Nine trials standardized or described at least 1 other cointervention such as sedation, 28,30,31 paralysis, 20–32 or ventilator weaning. 24,25,27,28,31,33 Four trials reported postrandomization withdrawals: less than 5% of enrolled patients in 3 trials (9/802, 27 1/102, 28 6/142 33) and 8% in 1 trial (2/24). Only 1 trial 27 reported any losses to follow-up (< 0.25% of patients or 2/802). However, only 2 of the 5 trials reporting ventilator-associated pneumonia 24,25,27,30,31 partially blinded outcome assessors to treatment group or adjudicated the outcome, 27 and only 3 trials provided specific diagnostic criteria for ventilator-associated pneumonia. 24,27,30 One trial classified pressure ulcers using standardized criteria. 23 Five trials 23,24,27,28,31 reported crossovers between groups, which involved less than 10% of patients in 4 trials (12/304, 24 2/21, 27 4/102, 28 5/142 33). All trials analyzed data for patients by assigned group. 22–34 Five trials ended early, 1 after meeting prespecified futility criteria, 29 and 4 because of low or declining enrolment. 23,24,31,34

Clinical outcomes

Mortality

In the primary analysis (10 trials, n = 1486), 22–24,27–32,34 ventilation in the prone position had no effect on mortality (RR 0.96, 95% CI 0.84 to 1.09; p = 0.52) (Figure 2). Three trials provided no mortality data. 25,26,34 In the subgroup analysis, we found no significant difference in mortality between trials of short-term prone positioning 22,29,33 (RR 0.77, 95% CI 0.46 to 1.28) and those of prolonged prone positioning 23,24,27,28,30,31,34 (RR 0.97, 95% CI 0.85 to 1.11; p = 0.39 for comparison of RRs

| Study | Prone, n/N | Supine, n/N | RR (95% CI) |
|-------|-----------|-------------|-------------|
| **Short-term prone positioning** | | | |
| Leal et al 22 | 5/8 | 6/8 | 0.83 (0.43-1.63) |
| Papazian et al 29 | 3/13 | 5/13 | 0.60 (0.18-2.01) |
| Dennery et al 32 | 4/13 | 6/15 | 0.77 (0.28-2.14) |
| **Subtotal** | 34 | 36 | 0.77 (0.46-1.28) |
| **Overall effect p = 0.32. Heterogeneity I^2 = 0%** | | | |

**Prolonged prone positioning**

| Study | Prone, n/N | Supine, n/N | RR (95% CI) |
|-------|-----------|-------------|-------------|
| Gattinoni et al 23 | 77/152 | 73/152 | 1.05 (0.84-1.32) |
| Beuret et al 24 | 3/12 | 4/9 | 0.56 (0.17-1.91) |
| Guerin et al 27 | 134/413 | 119/378 | 1.03 (0.84-1.26) |
| Curley et al 28 | 4/51 | 4/51 | 1.00 (0.26-3.78) |
| Voggenreiter et al 30 | 1/21 | 3/19 | 0.30 (0.03-2.66) |
| Mancebo et al 31 | 33/76 | 35/60 | 0.74 (0.53-1.04) |
| Chan et al 34 | 4/11 | 4/11 | 1.00 (0.33-3.02) |
| **Subtotal** | 736 | 680 | 0.97 (0.85-1.11) |
| **Overall effect p = 0.68. Heterogeneity I^2 = 0%** | | | |
| **Overall** | 770 | 716 | 0.96 (0.84-1.09) |

**Figure 2:** Effect of ventilation in the prone position on mortality. We used a random-effects model in our analysis. The duration of prone positioning was up to 24 hours for 1–2 days in the short-term trials and up to 24 hours daily for more than 2 days in the prolonged-duration trials. One trial included data only for patients with acute hypoxemic respiratory failure. Including all patients from this trial (7/25 deaths in the prone group and 14/28 deaths in the supine group) did not change the result (RR 0.95, 95% CI 0.85 to 1.08; p = 0.41). I^2 = percentage of total variation across studies owing to between-study heterogeneity rather than chance. CI = confidence interval, RR = risk ratio.
using z score) (Figure 2). However, a single trial investigating an extended proning strategy (mean of 17 hours per day for 10 days) for acute respiratory distress syndrome showed a trend toward reduced mortality (RR 0.74, 95% CI 0.53 to 1.04; \( p = 0.08 \)). The effect of prone positioning on mortality was also similar (nonsignificant z scores for differences in RRs) in 9 trials \((n = 681)\) among patients with acute lung injury or acute respiratory distress syndrome\(^{22–24,28–32,34}\) (RR 0.92, 95% CI 0.78 to 1.09; \( p = 0.35 \)) and in 6 trials \((n = 611)\) in the subset of these patients who had prolonged duration of prone positioning (RR 0.94, 95% CI 0.79 to 1.13; \( p = 0.53 \)). Additional subgroup analyses to explore the effects of study quality and age (children v. adults) were limited because, in each case, one of the subgroup pairs included only a single randomized controlled trial (Appendix 3, available at www.cmaj.ca/cgi/content/full/178/9/1153/DC2). All mortality analyses showed little to no statistical heterogeneity beyond that expected by chance (\( I^2 < 10\% \)).

Visual inspection of a funnel plot revealed asymmetry, which suggested under-reporting of smaller trials showing excess mortality in the prone group. However, results of statistical tests did not confirm publication bias (Appendix 4, available at www.cmaj.ca/cgi/content/full/178/9/1153/DC2). All mortality analyses showed no increased risks (Figure 4, Table 2). When data were recorded in the morning, just before the subsequent proning manoeuvre, when patients in the prone group were in the supine position. On day 1, patients in the prone group had already completed 1 manoeuvre.

### Ventilator-associated pneumonia and duration of ventilation

In 6 trials \((n = 1026)\), ventilation in the prone position reduced the risk of ventilator-associated pneumonia (RR 0.81, 95% CI 0.66 to 0.99, \( p = 0.04 \)), with no statistical heterogeneity beyond that expected by chance (\( I^2 = 0\%)\) (Figure 4). One of the 6 trials \((n = 22)\) recorded no cases of ventilator-associated pneumonia.\(^{34}\) The effect was similar in trials that blinded ventilator-associated pneumonia assessors\(^{34}\) or adjudicated the outcome\(^{27} \) (\( p = 0.89 \) for comparison of RRs in these v. other trials). Although in 6 trials \((n = 992)\) patients who received ventilation in the prone position generally had shorter durations of ventilation (weighted mean difference \(-0.9 \) days, 95% CI \(-1.9\) to 0.1; \( p = 0.06 \), \( F = 3\%)\), in 4 trials \((n = 148)\) the number of ventilator-free days in the prone group was not significantly greater than the number in the supine group (weighted mean difference 3.7 days, 95% CI \(-1.8\) to 9.3; \( p = 0.19 \), \( F = 67\% \)).

### Adverse events

In 6 trials \((n = 504)\), ventilation in the prone position increased the risk of pressure ulcers (RR 1.36, 95% CI 1.07 to 1.71; \( p = 0.01 \), \( F = 0\%)\). Between 5 and 8 trials contributed data to analyses of other adverse events; these analyses showed no increased risks (Figure 5, Table 2). When data

### Table 1: Effect of ventilation in the prone position on oxygenation, by timing of measurement of ratio of partial pressure of oxygen to inspired fraction of oxygen

| Timing of oxygenation measurement; day since randomization | No. of trials (patients) | Ratio of means† (95% CI) | \( p \) value | \( I^2 \), %‡ |
|----------------------------------------------------------|--------------------------|--------------------------|--------------|--------------|
| 1 hr after start of proning manoeuvre                     |                          |                          |              |              |
| Day 1                                                    | 4 (434)                  | 1.31 (1.12–1.53)         | < 0.001      | 76           |
| Day 2                                                    | 3 (379)                  | 1.25 (1.09–1.43)         | 0.001        | 55           |
| Day 3                                                    | 3 (330)                  | 1.24 (1.05–1.46)         | 0.01         | 68           |
| At the end of proning manoeuvre                          |                          |                          |              |              |
| Day 1                                                    | 8 (633)                  | 1.34 (1.23–1.45)         | < 0.001      | 29           |
| Day 2                                                    | 4 (379)                  | 1.30 (1.15–1.46)         | < 0.001      | 42           |
| Day 3                                                    | 5 (445)                  | 1.23 (1.15–1.32)         | < 0.001      | 0            |
| Just before subsequent proning manoeuvre§               |                          |                          |              |              |
| Day 1                                                    | 9 (1206)                 | 1.06 (1.00–1.12)         | 0.07         | 19           |
| Day 2                                                    | 7 (1106)                 | 1.09 (1.04–1.14)         | < 0.001      | 0            |
| Day 3                                                    | 6 (1045)                 | 1.09 (1.04–1.14)         | < 0.001      | 0            |

Note: CI = confidence interval.

*The author of 1 trial\(^{15}\) confirmed that the published error terms for ratios of partial pressure of oxygen to inspired fraction of oxygen were standard deviations rather than standard errors of means. Meta-analyses of oxygenation index (defined as 100 x mean airway pressure/ratio of partial pressure of oxygen to inspired fraction of oxygen) are not shown because only 3 trials\(^{28–30,31}\) reported these data at any time, with only 1 trial\(^{28}\) providing data beyond day 1. These 3 trials all reported the ratio of partial pressure of oxygen to inspired fraction of oxygen and oxygenation index simultaneously.

†Ratio of means is the mean ratio of partial pressure of oxygen to inspired fraction of oxygen in the prone group divided by that in the control group. Random-effects models were used in all analyses.

‡\( I^2 \) = percentage of total variation across studies owing to between-study heterogeneity rather than chance.

§Data were recorded in the morning, just before the subsequent proning manoeuvre, when patients in the prone group were in the supine position. On day 1, patients in the prone group had already completed 1 manoeuvre.
Interpretation

Our systematic review suggests that mechanical ventilation in the prone position does not improve survival for patients with acute hypoxemic respiratory failure, including acute lung injury and acute respiratory distress syndrome, despite improved oxygenation and a reduced risk of ventilator-associated pneumonia. Prone positioning increased the risk of endotracheal tube obstruction, but otherwise it was safe. Despite variable duration of ventilation in the prone position and clinically diverse populations in the included trials, pooled clinical outcomes had little statistical heterogeneity, which strengthens our findings. These results do not justify the routine use of prone positioning during mechanical ventilation in patients with acute hypoxemic respiratory failure.

There are several hypotheses that may explain the neutral effect of ventilation in the prone position on mortality. First, short- to medium-term improved oxygenation may not increase survival because of poor correlation between oxygenation and severity of lung injury. Indeed, patients with acute respiratory distress syndrome die more often of multiple organ failure than hypoxemia. Although prone positioning improves oxygenation within 1 hour and to a greater extent than inhaled nitric oxide, such improvements may help only the most severely hypoxic patients to survive. Gattinoni and colleagues reported a post hoc analysis showing a significantly lower 10-day mortality rate with prone positioning in the subgroup of patients with the lowest ratio of partial pressure of oxygen to inspired fraction of oxygen. In contrast, Mancebo and colleagues did not find the initial ratio of partial pressure of oxygen to inspired fraction of oxygen to be a significant predictor of mortality in a multivariable logistic regression model. Neither study reported a statistical test of the interaction between treatment group and initial ratio of partial pressure of oxygen to inspired fraction of oxygen. No randomized controlled trials have investigated prone positioning during ventilation as rescue therapy for critical hypoxemia. Prone positioning in such patients may prevent imminent death and allow time for other treatments to help.

Second, the broad nature of selection criteria in the included trials and the variability in duration of ventilation in the prone position may explain the neutral effect of prone ventilation on mortality. Finally, there are several hypotheses that may explain the neutral effect of ventilation in the prone position on mortality. First, short- to medium-term improved oxygenation may not increase survival because of poor correlation between oxygenation and severity of lung injury. Indeed, patients with acute respiratory distress syndrome die more often of multiple organ failure than hypoxemia. Although prone positioning improves oxygenation within 1 hour and to a greater extent than inhaled nitric oxide, such improvements may help only the most severely hypoxic patients to survive. Gattinoni and colleagues reported a post hoc analysis showing a significantly lower 10-day mortality rate with prone positioning in the subgroup of patients with the lowest ratio of partial pressure of oxygen to inspired fraction of oxygen. In contrast, Mancebo and colleagues did not find the initial ratio of partial pressure of oxygen to inspired fraction of oxygen to be a significant predictor of mortality in a multivariable logistic regression model. Neither study reported a statistical test of the interaction between treatment group and initial ratio of partial pressure of oxygen to inspired fraction of oxygen. No randomized controlled trials have investigated prone positioning during ventilation as rescue therapy for critical hypoxemia. Prone positioning in such patients may prevent imminent death and allow time for other treatments to help.

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### Table 1: Comparison of Prone and Supine Groups

| Study                  | Prone, N | Supine, N | Ratio of means (95% CI) |
|------------------------|----------|-----------|-------------------------|
| **Day 1**              |          |           |                         |
| Gattinoni et al        | 147      | 148       | 1.28 (1.15–1.42)        |
| Watanabe et al         | 8        | 8         | 1.39 (1.16–1.66)        |
| Curley et al           | 45       | 49        | 1.14 (0.95–1.37)        |
| Papazian et al         | 13       | 13        | 1.64 (1.29–2.10)        |
| Mancebo et al          | 71       | 59        | 1.27 (1.09–1.49)        |
| Chan et al             | 11       | 11        | 1.53 (1.00–2.33)        |
| Demory et al           | 13       | 15        | 1.69 (1.28–2.24)        |
| Ibrahim et al          | 11       | 11        | 1.30 (1.02–1.66)        |
| Overall                | 319      | 314       | 1.34 (1.23–1.45)        |
| Overall effect p < 0.001. Heterogeneity I² = 29%. |
| **Day 2**              |          |           |                         |
| Gattinoni et al        | 121      | 139       | 1.25 (1.13–1.39)        |
| Watanabe et al         | 8        | 8         | 1.38 (1.16–1.65)        |
| Curley et al           | 41       | 47        | 1.19 (1.00–1.41)        |
| Chan et al             | 8        | 7         | 2.09 (1.26–3.47)        |
| Overall                | 178      | 201       | 1.30 (1.15–1.46)        |
| Overall effect p < 0.001. Heterogeneity I² = 42%. |
| **Day 3**              |          |           |                         |
| Gattinoni et al        | 95       | 132       | 1.20 (1.07–1.34)        |
| Watanabe et al         | 8        | 8         | 1.46 (1.21–1.76)        |
| Curley et al           | 29       | 41        | 1.16 (0.96–1.41)        |
| Mancebo et al          | 65       | 52        | 1.22 (1.06–1.40)        |
| Chan et al             | 8        | 7         | 1.08 (0.66–1.77)        |
| Overall                | 205      | 240       | 1.23 (1.15–1.32)        |
| Overall effect p < 0.001. Heterogeneity I² = 0%. |

**Figure 3:** Effect of ventilation in the prone position on daily ratio of partial pressure of oxygen to inspired fraction of oxygen. We used a random-effects model in our analysis. Values were recorded at the end of the period of prone positioning (prone group) and simultaneously in the supine group. Ratio of means = mean ratio of partial pressure of oxygen to inspired fraction of oxygen in the prone group divided by that in the supine group. CI = confidence interval.
cluded trials may have failed to identify a particular population that would benefit from ventilation in the prone position. Several trials enrolled patients with diverse types of respiratory failure. In the largest trial, only 51% of the 802 patients had acute lung injury or acute respiratory distress syndrome. Although meta-analysis restricted to patients with acute lung injury or acute respiratory distress syndrome did not show a mortality benefit, physiologic variables (other than oxygenation response) may identify a subgroup of these patients who might benefit from ventilation in the prone position. For example, Gattinoni and colleagues reported in a post hoc analysis that decreased partial pressure of carbon dioxide after an initial 6-hour prone period, which likely reflected a lower fraction of minute ventilation delivered to nonperfused lungs (dead space), was associated with improved survival.

Third, the duration of prone positioning may have been insufficient. Our subgroup analysis did not show benefit among all prolonged-duration trials; however, the most intensive proning regimen studied (17 hours daily for 10 days) was shown in an adjusted analysis to reduce mortality. An ongoing randomized controlled trial of mechanical ventilation with an intensive proning regimen (20 hours daily for up to 28 days) involving 340 patients with acute respiratory distress syndrome (the Prone–Supine Study II [www.clinicaltrials.gov/ct2/show/NCT00159939?term=NCT00159939&rank=1]) may provide valuable additional data regarding this issue.

Finally, it is possible that the benefits of ventilation in the prone position were overshadowed by a mechanical ventilation strategy that injured the lungs and perpetuated multiple organ failure. The ongoing Prone–Supine Study II mandates lung-protective mechanical ventilation, which may resolve this issue. Another planned randomized controlled trial would add to current knowledge by enrolling patients with severe acute respiratory distress syndrome (ratio of partial pressure of oxygen to inspired fraction of oxygen no greater than 100 mm Hg and high severity of illness) and mandating prolonged prone positioning and lung-protective ventilation.

Our meta-analysis demonstrated a reduced risk of ventilator-associated pneumonia associated with prone positioning. Some small nonrandomized studies suggested better drainage of respiratory secretions with this technique, which may prevent aspiration. However, our finding is limited by potential ascertainment bias because most of the trials lacked standard diagnostic criteria and blinding of outcomes assessors. Furthermore, the clinical importance of reduced ventilator-associated pneumonia with prone positioning is

**Table 1**

| Study            | Prone, n/N | Supine, n/N | RR (95% CI) |
|------------------|------------|-------------|-------------|
| Beuret et al     | 2/12       | 4/9         | 0.38 (0.09–1.62) |
| Watanabe et al   | 1/8        | 2/8         | 0.50 (0.06–4.47) |
| Guerin et al     | 85/413     | 91/378      | 0.85 (0.66–1.11) |
| Voggenreiter et al | 13/21     | 17/19       | 0.69 (0.48–1.00) |
| Mancebo et al    | 14/76      | 9/60        | 1.23 (0.57–2.64) |
| Chan et al       | 0/11       | 0/11        | NA          |
| Overall          | 541        | 485         | 0.81 (0.66–0.99) |

**Figure 4**: Effect of ventilation in the prone position on risk of ventilator-associated pneumonia. We used a random-effects model in our analysis. One trial included data only for patients with acute hypoxemic respiratory failure. $I^2$ = percentage of total variation across studies owing to between-study heterogeneity rather than chance. CI = confidence interval, NA = not applicable, RR = risk ratio.

**Table 2**

| Study            | Prone, n/N | Supine, n/N | RR (95% CI) |
|------------------|------------|-------------|-------------|
| Leal et al       | 1/8        | 0/8         | 3.00 (0.14–64.26) |
| Gattinoni et al  | 55/152     | 42/152      | 1.31 (0.94–1.83) |
| Beuret et al     | 2/12       | 2/9         | 0.75 (0.13–4.36) |
| Curley et al     | 10/51      | 8/50        | 1.23 (0.53–2.85) |
| Voggenreiter et al | 19/21    | 12/19       | 1.43 (0.99–2.07) |
| Chan et al       | 2/11       | 0/11        | 5.00 (0.27–93.55) |
| Overall          | 255        | 249         | 1.36 (1.07–1.71) |

**Figure 5**: Effect of ventilation in the prone position on risk of pressure ulcers. We used a random-effects model in our analysis. One trial included data only for patients with acute hypoxemic respiratory failure. $I^2$ = percentage of total variation across studies owing to between-study heterogeneity rather than chance. CI = confidence interval, RR = risk ratio.
unclear given similar duration of ventilation, ventilator-free days and mortality among patients who received ventilation in the prone and supine positions.

We found mechanical ventilation in the prone position to be generally safe, at least in centres participating in clinical trials. However, the procedure was labour-intensive, requiring 4–6 personnel to turn an adult patient, and it increased the risk of pressure ulcers. In contrast, with the exception of possible endotracheal tube blockage, prone positioning did not increase the risk of potentially life-threatening complications such as accidental extubation and dislodgement of central catheters or thorocostomy tubes. Nevertheless, some trials reported airway obstruction, accidental extubation and dislodgement of central catheters directly related to the turning procedure,23,31 which led to cardiac arrest in one instance.31 Such complications, although infrequent, could be catastrophic in patients with critical hypoxemia. Less experienced centres may face more life-threatening complications, but turning protocols and nursing care guidelines may mitigate the risk.33,36,47

A recent survey of 702 (predominantly adult) ICUs in Germany46 reported more complications during various forms of positioning therapy than recorded in randomized controlled trials. These included hemodynamic instability (reported by 74% of ICUs), accidental removal of tubes or catheters (50%), worsening gas exchange (45%), patient intolerance owing to inadequate sedation (41%) and cardiac arrhythmias (22%). Moreover, many respondents believed that successful application of positioning therapy forced clinicians to compromise other aspects of critical care by requiring them to deepen sedation (77% of ICUs) and to stop (16%) or at least reduce (33%) enteral feeds. Another survey of proning practices in 25 ICUs in Belgium reported similar rates of serious complications, in addition to increased workload (owing to increased suctioning and eye care), which contributed to the reluctance of nurses in 9 ICUs (36%) to use the technique.22

Strengths of our review include methods to reduce bias and analysis of a comprehensive set of prespecified clinical and physiologic outcomes. Our study had a number of limitations, including variability in the selection criteria of individual trials (including author definitions of acute lung injury and acute respiratory distress syndrome) and reduced sample size in 5 trials that ended early because of futility or declining enrolment. Both factors may have diluted our ability to detect a survival benefit through meta-analysis. In addition, although the mortality funnel plot suggested publication bias (Appendix 4, available at www.cmaj.ca/cgi/content/full/178/9/1153/DC2), results of statistical tests did not confirm the presence of publication bias. Statistical tests may fail to detect publication bias, but the assumption of such bias in our meta-analysis would imply that small unpublished randomized trials have shown higher rates of mortality in the prone group than the supine group. Data from such trials would only move the estimated pooled RR for mortality closer to no effect. Another limitation is that supplementary information was not available for all trials, 2 of which were published only as abstracts. In addition, the small number of trials included in our review reduced the precision of the pooled estimates for some clinical and physiologic analyses and may have underestimated heterogeneity. Finally, findings from the largest trial27 dominated the meta-analysis of ventilator-associated pneumonia and endotracheal tube obstruction.

In summary, our systematic review found that ventilation in the prone position in patients with acute hypoxemic respiratory failure improved oxygenation and reduced the risk of ventilator-associated pneumonia, but it did not improve survival. The technique appeared safe in expert centres; however, serious airway, catheter and tube complications may occasionally occur, and the technique increased the risk of pressure ulcers. Consequently, we do not recommend the routine use of prone positioning for patients with hypoxemic respiratory failure. Despite the neutral effect on mortality, clinicians may still consider

| Table 2: Risk of adverse events in 12 trials of prone positioning for mechanical ventilation included in our systematic review* |
|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|
| Adverse event   | No. of trials (patients) contributing data† | No. (%) of patients with adverse event‡ | RR (95% CI)        | p value | I², %§ |
| Pressure ulcers* | 6 (504) | 153/504 (30.4) | 1.36 (1.07-1.71) | 0.01  | 0     |
| Endotracheal tube obstruction* | 5 (204) | 3/204 (1.5) | 1.32 (0.09-18.50) | 0.84  | 33    |
| Accidental extubation* | 8 (662) | 44/662 (6.6) | 0.88 (0.48-1.60) | 0.67  | 0     |
| Loss of central venous or arterial access | 7 (526) | 25/526 (4.8) | 0.67 (0.31-1.44) | 0.31  | 0     |
| Thorocostomy tube dislodgement | 6 (504) | 7/504 (1.4) | 6.00 (0.73-49.24) | 0.10  | NA    |
| Pneumothorax* | 6 (336) | 16/336 (4.8) | 0.93 (0.35-2.45) | 0.89  | 0     |
| Cardiac arrest* | 5 (230) | 0/230 (0) | NA | NA | NA |

Note: CI = confidence interval, NA = not applicable, RR = risk ratio. Random-effects models were used in all analyses.

*We excluded the trial by Guerin et al.27 from the analysis because it reported the number of occurrences of adverse events rather than the number of patients with adverse events. Assuming a similar distribution of occurrences per patient in the prone and supine groups, when we included these data, the risk of endotracheal tube obstruction became statistically significant (RR 2.46, 95% CI 1.33-4.55; p = 0.004; I² = 0%) and the risk of pressure ulcers became more significant (RR 1.25, 95% CI 1.10-1.43; p = 0.001; I² = 0%). This trial contributes heavily to the pooled RRs when its data are included (92% weighting in the endotracheal tube obstruction analysis and 70% in the pressure ulcer analysis).

†We counted trials that collected data on the adverse event, regardless of whether an event occurred.

‡This is the unweighted proportion of patients in the prone and supine groups (in trials collecting data on adverse events) who experienced an adverse event.

§% = percentage of total variation across studies owing to between-study heterogeneity rather than chance.
prone positioning for life-threatening hypoxemia, along with other supportive therapies. Current data for such patients are limited, and early termination of several published trials owing to slow enrolment suggests that additional studies, although highly desirable, will be challenging to complete.

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