Perioperative Allogenenic Blood Transfusion Is Associated with Worse Clinical Outcomes for Hepatocellular Carcinoma: A Meta-Analysis

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Abstract

Background and Objective: The impact of perioperative allogenenic blood transfusion (ABT) on clinical outcomes for hepatocellular carcinoma (HCC) is conflicting and unclear. The aim of this meta-analysis is to evaluate the association between ABT and HCC clinical outcomes. Outcomes evaluated were all-cause death, tumor recurrence and postoperative complications.

Methods: Relevant articles were identified through MEDLINE search (up to November 2012). Meta-analyses were performed by using the fixed or random effect models. Study heterogeneity was assessed by Q-test and I² test. Publication bias was evaluated by funnel plots, Egger’s and Begg’s test.

Results: A total of 5635 cases from 22 studies finally met our inclusion criteria. Meta-analysis indicated HCC patients with ABT had an increased risk of all-cause death at 3 and 5 years after surgery (respectively: OR = 1.92, 95% CI, 1.61–2.29, P < 0.001; OR = 1.60, 95% CI, 1.47–1.73, P < 0.001) compared with those without ABT. The risk of tumor recurrence was significantly higher for ABT cases at 1, 3 and 5 years (respectively: OR = 1.70, 95% CI, 1.38–2.10, P < 0.001; OR = 1.22, 95% CI, 1.08–1.38, P < 0.001; OR = 1.16, 95% CI, 1.08–1.24, P < 0.001). The HCC cases with ABT significantly increased postoperative complications occurrence compared with non-ABT cases (OR = 1.78, 95% CI, 1.34–2.37, P < 0.001).

Conclusions: The findings from the current meta-analysis demonstrated that ABT was associated with adverse clinical outcomes for HCC patients undergoing surgery, including increased death, recurrence and complications. Therefore, ABT should not be performed if possible.

Introduction

Hepatocellular carcinoma (HCC) is one of the most common malignancies all over the world, especially in Asia and Africa [1]. Although liver transplantation has already been proven as an alternative treatment method for HCC during the early stage, the scarcity of donors limits this treatment [2]. Hepatectomies is still the mainstay of treatment for patients with HCC. Most patients with HCC have poor hepatic function reserve. What is more, patients with liver cirrhosis showed decreased platelet count, which results in an increased risk of hemorrhage and prothrombin activity during surgery. Over the past decades, surgical techniques, perioperative care and increased experience have improved the safety of liver resection for HCC [3]. Despite these advances, partial liver resection still carries the risk for excessive blood loss and a subsequent need for blood transfusion. Blood transfusion rate during hepatic resections have been reported to be decreased from 62% in 1985 to 22% in 2007 [4].

Blood transfusion could save patients’ life once substantial hemorrhage occurred. The type of blood transfusion could be divided into autotransfusion and allogenenic blood transfusion (ABT). Autotransfusion represents autologous collection and reinfusion of patients’ own blood or blood components before surgery, while ABT refers to receive blood from other persons. However, ABT has been reported to be associated with potentially devastating complications, such as transmission of human immuno-deficiency virus and hepatitis, transfusion reactions and increased postoperative infection rate [5]. Noninfectious risks such as transfusion-associated circulatory overload and transfusion-related acute lung injury (TRALI) are also well known [6,7]. There are several studies focused on the long-term outcomes after ABT in patients undergoing curative resection of HCC. Hanazaki et al [8]...
have reported that there were adverse effects of ABT on cancer recurrence and survival rate for HCC. Whereas, other investigators gave the inconsistent results. Kwon and his colleagues [9] demonstrated that ABT has no influence on the long-term oncological outcome for HCC.

It has been seen that the association between ABT and clinical outcomes for HCC remains controversial. The published results were largely based on a retrospective analysis of cases from a single center, which could contribute to this inconsistent phenomenon. The method of meta-analysis could fully solve this problem, which played an important role in Evidence-medicine [10]. The rationale for a meta-analysis is that, by combining the samples of individual studies, the overall sample size is increased, thereby improving the statistical power of the analysis as well as the precision of the estimates of treatment effects [11]. As a result, on the basis of relevant literature search, we performed a comprehensive meta-analysis to fully estimate the postoperative influence of ABT for patients with HCC.

Methods

This meta-analysis was done in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-analyses (PRISMA) guideline [12].

Identification and Eligibility of Relevant Studies

Studies were identified via a systematical search of the electronic database Pubmed (since 1967). The last search was performed on November 30th, 2012. The following keywords were used: “Hepatocellular carcinoma (liver cancer)” and “blood transfusion”. And our search was restricted to articles published in English. Relevant lists of all relevant publications were hand searched for additional studies missed by the search strategy. To be included in our meta-analysis, the papers should meet all of the following criteria: (1) The studies had to evaluate the association between ABT and clinical outcomes (postoperative complications, recurrence or death) for HCC. (2) The report contained original data. (3) The papers focused on non-primary hepatocellular carcinoma (e.g. benign tumors and metastatic neoplasm) or autogenic blood transfusion were excluded. (4) Reviews and letters to editors were not included because of insufficient data for analysis. If the same population was examined in the multiple studies, only the most recent report was included in the analysis.

Outcomes

The primary outcome measures were postoperative complications, postoperative occurrences of tumor recurrence and all-cause death. The predefined time points for recurrence and all-cause death was three-year and five-year after surgery. Other time points such as one-year or ten-year could also be taken into account. The cause of death, if available, was analyzed for assessment of cancer-related mortality.

Data Abstraction and Quality Assessment

In our study, we divided all cases into ABT group and non-ABT group. The ABT group was defined as cases who received any amount of allogeneic blood products (packed red blood cells, plasma, platelets or whole blood). And the non-ABT group was defined as cases who did not receive any blood products at all. The perioperative period included 30 days before and after surgery. Data abstraction was independently run by two of the authors (Liu and Wang). If they generated different results, a third investigator (Zhang) was asked to discuss to reach an agreement. Information retrieved from the original reports included first author’s name, publication year, study design, patients’ age and sex, liver cirrhosis, Child-Pugh class, tumor stage, amount of blood transfusion,
| Author                       | Publication Year | Study type | Number of patients | ABT Rate (%) | Age (mean, Y) | Sex (male%) | Liver Cirrhosis (%) | Child A(%) | TNM stage | Transfused Units | Study quality |
|------------------------------|------------------|------------|--------------------|--------------|---------------|-------------|---------------------|------------|------------|-----------------|---------------|
| Matsumata et al16            | 1993             | RCS        | 54 72 126          | 42.86        | 57.6          | 83.3        | 56.5                |            |            |                 | 9             |
| Itasaka et al17              | 1995             | PCS        | 38 33 71           | 53.52        | 58.1          | 85.9        | 90.1                | 94.4       |            |                 | 9             |
| Yamamoto et al18             | 1996             | RCS        | 85 301 386         | 22.02        | 59.7          | 80.6        | 59.3                | 68.7       |            |                 | 6             |
| Asahara et al19              | 1999             | RCS        | 23 152 175         | 13.14        | 60.1          | 76.0        | 70.3                | 75         | 56         |                 | 8             |
| Makino et al20               | 2000             | RCS        | 117 78 195         | 60.00        | 60.8          | 80.0        | 71.3                | 67.7       |            |                 | 9             |
| Kitagawa et al21             | 2001             | PCS        | 26 52 78           | 33.33        | 60.0          | 55.5        |                     |            |            |                 | 7             |
| Kwon et al9                  | 2001             | PCS        | 53 55 108          | 49.07        | 62.0          | 76.9        | 48.2                | 80.0       | 108        | 0 5.3          | 9             |
| Ercolani et al22             | 2002             | RCS        | 71 153 224         | 31.07        | 62.5          | 81.3        | 78.6                | 82.6       |            |                 | 6             |
| Sasaki et al23               | 2002             | PCS        | 23 52 75           | 30.27        | 61.7          | 78.7        | 61.3                | 55         | 20         |                 | 5             |
| Wei et al24                  | 2003             | RCS        | 45 110 155         | 29.03        | 52.0          | 78.7        | 34.9                |            |            |                 | 3.5           |
| Hanazaki et al8              | 2005             | RCS        | 210 158 368        | 57.07        | 62.4          | 75.6        | 52.2                | 78.3       | 300        | 68             | 9             |
| Laurent et al25              | 2005             | RCS        | 36 72 108          | 33.33        | 64.0          | 82.4        | 63.0                | 68         | 40         |                 | 6             |
| Sasaki et al26               | 2006             | RCS        | 184 233 417        | 44.12        | 74.2          | 67.9        | 74.2                | 339        | 78         |                 | 8             |
| Sugita et al27               | 2008             | RCS        | 101 120 221        | 45.70        | 63.2          | 77.0        | 54.5                |            |            |                 | 7.9           |
| Kaibori et al28              | 2008             | RCS        | 269 141 410        | 56.61        | 64.2          | 79.8        | 39.5                | 264        | 136        | 5.0            | 9             |
| Wang et al29                 | 2009             | RCS        | 62 411 473         | 13.11        | 53.1          | 80.1        | 44.4                | 83.3       | 400        | 73             | 8             |
| Choi et al30                 | 2009             | RCS        | 91 96 187          | 48.66        | 51.0          | 75.4        | 56.7                |            |            |                 | 9             |
| Abdel-Wahab et al31          | 2010             | RCS        | 87 72 159          | 54.72        | 54.8          | 75.0        | 87.0                | 138        | 37         | 3.0            | 7             |
| Nanashima et al32            | 2011             | RCS        | 100 83 183         | 54.64        | 65.0          | 80.9        |                     |            |            |                 | 7             |
| Okamura et al33              | 2011             | RCS        | 87 289 376         | 23.14        | 61.4          | 80.6        | 94.2                | 213        | 163        |                 | 7             |
| Yang et al34                 | 2011             | RCS        | 164 141 305        | 53.77        | 49.2          | 91.5        | 76.7                | 88.5       | 117        | 188            | 5.8           |
| Kuo et al35                  | 2012             | RCS        | 76 759 835         | 9.10         | 63.5          | 76.9        | 516                  | 85.4       |            |                 | 4.2           |

ABT+, allogenic blood transfused; ABT−, non-allogenic blood transfused; RCS, retrospective cohort study; PCS, prospective cohort study.

*Study provided cancer-related death data and could not be included.
doi:10.1371/journal.pone.0064261.t001
Figure 2. Forest plot of postoperative death risk associated with ABT for HCC. (A) for 3-year and (B) for 5-year. OR, odds ratio; CI, confidence interval.

doi:10.1371/journal.pone.0064261.g002
Meanwhile, I2-test was also utilized to examine the heterogeneity fixed-effect model (the Mantel-Haenszle method) was applied [14]. The random-effect model (the DerSimonian and Laird method) was used to calculate summary estimate; otherwise, the studies were rated as poor quality. The studies were defined as a study with a quality score more than 7 points, otherwise the studies were rated as poor quality.

Statistical Analysis

For each outcome, odds ratio (OR) and its 95% confidence intervals (CI) was used to measure the association for each study. In this analysis, we examined the possible heterogeneity between studies using the $X^2$-based Q test [13]. Where there was evidence of heterogeneity between studies (the P value for Q test less than 0.05), the random-effect model (the DerSimonian and Laird method) was used to calculate summary estimate; other wise, the fixed-effect model (the Mantel-Haenszle method) was applied [14]. Meanwhile, $I^2$-test was also utilized to examine the heterogeneity (values of 25%, 50% and 75% were considered to represent low, medium, and high heterogeneity respectively) [15]. The potential source of heterogeneity was evaluated by meta-regression (publication year, study design, sample size and sample constitution (the proportion of patients who underwent blood transfusion in each study)). Stratified analysis centered on potential confounders was also performed to evaluate the robustness of the conclusion drawn. Additionally, sensitivity analyses were carried out by excluding the highest or lowest OR value to assess the stability of the association between ABT and cancer clinical outcomes. Publication bias was assessed by funnel plots [16]. And the symmetry of funnel plots was analyzed using Egger’s and Begg’s test. All analyses were conducted using STATA 11.0 statistical software (Stata Corp, College Station, TX, USA).

### Results

#### Literature Search and Study Characteristics

Figure 1 shows a flow diagram of how we selected relevant articles. The abstract of 177 primary studies were identified for initial review using the above-mentioned search strategies. After screening the abstracts, 58 full-text articles were retrieved for more detailed assess. Meanwhile, the hand search identified a further 5 articles. Eight were eliminated because of published language not in English, and thirty-one studies were also excluded due to duplication of data or inadequate data for analysis. Only one study reported the cancer-related survival data, it was impossible for us to adopt it to perform survival analysis. In total, 22 studies[8;9;17;18;19;20;21;22;23;24;25;26;27;28;29;30;31;32;33;34;35;36] met our inclusion criteria and were included in the final meta-analysis. Study characteristics were listed in Table 1.

All included studies gave a total of 5635 patients, of whom 2002 (35.5%) cases received ABT and 3633 (64.5%) cases were grouped as non-ABT. The sample size for each study varied from 71 to 835 (median: 175). The reported ABT rate in each article ranged from 9.10% to 65.61%. The average of all patients in all included studies was 61 years. The total proportion of male subjects was around 80% in all included studies. Of the 22 studies, 18 reported the presence or absence of liver cirrhosis, with approximately 60% of patients having liver cirrhosis. Twenty-two eligible studies comprised 4 prospective studies and 18 retrospective studies. Outcomes reported in each article included all-cause death (n = 12), tumor recurrence (n = 16) and postoperative complications (n = 7). All included studies were published between 1993 and 2012, of which 60% (n = 14) were published in 2007 or more recent years. According to the 9-star Newcastle-Ottawa Scale, half (n = 11) of the studies were defined as high-quality studies (score more than 7).

#### All-cause Death

Twelve studies reported the all-cause death, giving a total sample size of 3113 patients for evaluation. Of them, 4 studies[8;22;29;36] provided postoperative 3-year survival data
and 12 studies [8;9;17;20;22;26;29;30;31;32;33;34;35;36] provided 5-year survival data. As shown in Figure 2A, meta-analysis data showed a significant higher death risk for patients with ABT than those without ABT during the 3 years after surgery, corresponding to a pooled OR of 1.92 (95% CI: 1.54–1.90, P<0.001). Between-study heterogeneity was not observed (I² = 19.0%, P = 0.295). Regarding to postoperative 5-year death risk, meta-analysis of all studies demonstrated a significant increased all-cause death after ABT, with an overall OR of 1.60 (95% CI: 1.47–1.73, P<0.001) (Figure 2B). Significant heterogeneity among studies was not present (I² = 18.8%, P = 0.259).

Table 2 shows the results of the stratified meta-analysis for postoperative 5-year all-cause death. When stratified by sample size, ABT could increase the postoperative all-cause death risk in large sample-size studies with the overall estimate (OR = 1.65, 95% CI: 1.49–1.83, P<0.001). Similar results were obtained in small sample-size studies (OR = 1.48, 95% CI: 1.30–1.69, P<0.001). By sample constitution, ABT was associated with increased risk of death in low ABT rate group (OR = 1.36, 95% CI: 1.38–1.62, P<0.001) and high ABT rate group (OR = 1.60, 95% CI: 1.44–1.77, P<0.001). Studies with high quality (OR = 1.71, 95% CI: 1.51–1.93, P<0.001) demonstrated slightly increased death risk than those with poor evidence level (OR = 1.50, 95% CI: 1.29–1.73, P<0.001). When examining differences over time, we found that studies published before 2007 had a summary estimate with a pooled OR of 1.71 (95% CI: 1.54–1.90, P<0.001), while studies published after 2007 had an overall estimate with a pooled OR of 1.45 (95% CI: 1.23–1.72, P<0.001). Sensitivity analysis was done for postoperative 5-year all-cause death by excluding the studies with the highest and the lowest OR value [17;29] and analyzing the impact on the results. The result did not alter the overall effect, giving a pooled OR of 1.58 (95% CI: 1.45–1.72, P<0.001). Meanwhile, heterogeneity tests between subgroups were also performed. The results showed no heterogeneity was observed (P>0.05) (Table 2).

Tumor Recurrence

Sixteen studies reported the data of disease progression defined as recurrence in ABT and non-ABT patients, involving 4120 cases. Postoperative 1-year data were available for five studies [9;17;19;21;30], 3-year data were reported by eleven studies [8;9;17;19;21;22;23;24;30;33;36] and 5-year data were provided by sixteen studies [8;9;17;19;20;21;22;23;24;26;27;28;29;30;31;32;33;34;35;36]. As displayed in Figure 3A, the result of meta-analysis suggested a significantly increased risk of recurrence for patients with ABT during the one year after surgery (OR = 1.70, 95% CI: 1.38–2.10, P<0.001). Significant heterogeneity did not exist among these studies (I² = 37.9%, P = 0.169), hence the fixed effect module was adopted.

With regard to the 3-year risk after operation, the overall estimate also showed patients with ABT had a significantly increased risk of recurrence than those without ABT (OR = 1.22, 95% CI: 1.08–1.38, P<0.001) (Figure 3B). However, significant heterogeneity between studies was observed (I² = 61.5%, P = 0.004), and the random effect module was applied. As for the postoperative 5-year risk of recurrence, the meta-analysis indicated patients with ABT had a significantly increased risk compared with non-ABT patients, giving a summary OR of 1.16 (95% CI: 1.08–1.24, P<0.001) (Figure 3B). However, significant heterogeneity between studies was observed (I² = 55.8%, P = 0.005), and the random effect module was applied.

As for the postoperative 5-year risk of recurrence, the meta-analysis indicated patients with ABT had a significantly increased risk compared with non-ABT patients, giving a summary OR of 1.16 (95% CI: 1.08–1.24, P<0.001) (Figure 3C). The random effect module was implemented due to slightly significant heterogeneity across the studies (I² = 55.8%, P = 0.005). A meta-regression was performed to evaluate the potential source of heterogeneity, which regressed the parameters of publication year, sample size, study design and sample constitution. Unfortunately, we failed to detect any factor that was recognized as the main source of between-study heterogeneity (P>0.05) (data not shown). Sensitivity analysis was carried out by exclusion of the studies with the highest and the lowest OR value [8;35]. Nearly no changes of the pooled OR was observed (OR = 1.17, 95% CI: 1.12–1.22, P<0.001).

Figure 3. Forest plot of postoperative recurrence risk associated with ABT for HCC. (A) for 1-year, (B) for 3-year and (C) for 5-year. OR, odds ratio; CI, confidence interval. doi:10.1371/journal.pone.0064261.g003
Postoperative Complications

Seven studies examined the relationship between ABT and postoperative complications, which enrolled 2260 patients. The combined OR of postoperative complications based on these studies was 1.78 (95% CI: 1.34–2.37, \( P < 0.001 \)) (Figure 4). Significant heterogeneity among studies was present (\( I^2 = 79.1\% \), \( P < 0.001 \)). Because of limited number of studies, sensitivity analysis was not able to be carried out in our study.

Publication Bias

We used Egger’s test and Begg’s test to evaluate the publication bias. No indication of publication bias was observed for all-cause death at 3 year (Egger’s test, \( P = 0.876 \); Begg’s test, \( P = 1.000 \)), and at 5 year (Egger’s test, \( P = 0.183 \); Begg’s test, \( P = 0.115 \)); for tumor recurrence at 1 year (Egger’s test, \( P = 0.130 \); Begg’s test, \( P = 0.142 \)), at 3 year (Egger’s test, \( P = 0.682 \); Begg’s test, \( P = 0.533 \)), and at 5 year (Egger’s test, \( P = 0.590 \); Begg’s test, \( P = 0.533 \)).

Figure 4. Forest plot of postoperative complications risk associated with ABT for HCC. OR, odds ratio; CI, confidence interval.
doi:10.1371/journal.pone.0064261.g004

Figure 5. Funnel plot of publication bias test for ABT and HCC clinical outcomes. Postoperative death at 3 year (upper left), postoperative death at 5 year (upper middle), postoperative recurrence at 1 year (upper right), postoperative recurrence at 3 year (lower left), postoperative recurrence at 5 year (lower middle), postoperative complications (lower right). Horizontal axis represents the log of OR; vertical axis represents SE of log(OR); CI, confidence interval; OR, odds ratio; Log, logarithm; SE, standard error.
doi:10.1371/journal.pone.0064261.g005
95% CI: 1.41–1.97, \( P = 0.882 \); for postoperative complications (Egger’s test, \( P = 0.747 \); Begg’s test, \( P = 0.891 \)). The funnel plots for the relation between ABT and the clinical outcomes of HCC are shown in Figure 5.

Discussion

The possibility to store and transfuse blood has been a major advance in medicine in the 20th century, saving countless lives. Nevertheless, the relationship between ABT and tumor clinical outcomes has been debated in the literature for several decades. A number of authors have reported that ABT was associated with deleterious outcomes for HCC. For instance, in a series of 183 patients undergoing hepatectomy for HCC, Hanazaki et al [8] have demonstrated that ABT was a risk factor for tumor recurrence and worse prognosis. Whereas, Kwon and his colleagues [9] reported no effect of ABT on survival after hepatic resection. There is a lack of systematic evaluation focused on ABT for HCC. Therefore, we comprehensively searched through PubMed for all medical literature published in English-language journals until 2012 and performed a meta-analysis of the relationship between ABT and HCC clinical outcomes.

To our knowledge, this study provided the first comprehensive meta-analysis of effect of ABT on clinical outcomes after surgery for HCC. Twenty-two studies with a combined patient population of more than 5000 patients during the past two decades all over the world met the criteria and were included in the final meta-analysis. Our meta-analysis illustrated that ABT in patients undergoing surgery for HCC was associated with significantly worse outcomes, including all-cause death, tumor recurrence and postoperative complications. The meta-OR for all-cause death during the five years after surgery was 1.60 in this study. It means that patients with ABT had a 60% increased death risk than those without ABT. In terms of postoperative tumor recurrence, the combined OR value for five years was 1.16 in our study. That is to say, compared with patients without ABT, those with ABT had a more than 16% chance of tumor recurrence. Similar adverse effect of ABT on clinical outcomes was also observed in other malignancies. For instance, a recently published meta-analysis based on 20795 cases has proven that ABT was significantly associated with increased all-cause death (OR = 1.72, 95% CI: 1.55–1.91, \( P<0.001 \)) and more tumor recurrence (OR = 1.66, 95% CI: 1.41–1.97, \( P<0.001 \)) for colorectal cancer patients [37]. Meanwhile, this study suggested ABT was also associated with other adverse outcomes, such as postoperative infections, need for surgical reintervention and increased length of hospital stay. Another previous meta-analysis performed by Yao and his colleagues on ampullary cancer suggested that interoperative surgical reintervention and increased length of hospital stay. Meanwhile, this study suggested ABT was also associated with increased all-cause death (OR = 1.72, 95% CI: 1.55–1.91, \( P<0.001 \)) for colorectal cancer patients. Our study has other advantages except for the consistency of overall results. Up to now, the number of cases included in our study is one of the largest. Although publication bias is one of the major drawbacks of any meta-analysis, no publication bias was detected by Egger’s and Begg’s test, which indicated that statistical results approximated the truth. However, several limitations should be noted in our study. Firstly, all the included articles were non-randomized studies, and the evidence level is lower than that of randomized controlled trial (RCT). One reason for this is that there is no choice for patients with substantial hemorrhage but transfusion. Hence, this needs to be kept in mind when interpreting our result. In the absence of RCT now or in the future, this meta-analysis may represent a comprehensive analysis of the impact of ABT for patients with HCC. Secondly, the characteristics of patients may differ between the studies. As we know, the factors impacted the clinical outcomes of HCC patients after surgery include hepatitis C infection, pre-treatment α-fetoprotein level, liver cirrhosis and tumor stage[47;48]. Because of limited information, we were not able to control these potential confounded factors. Thirdly, surgery-related risk factors such as type and duration of operation should not be neglected. Postoperative transcathearteral chemoembolization (TACE) has been involved significantly over the past years, with significant improvement in survival[49;50]. The adjuvant treatment received by each patient was not fully understood, which could lead unknown effect to our results. Meanwhile, the lack of individual case data limited our ability to analyze the cancer-related death risk in our study. Future studies should prospectively record these data. Last but not least, although there was no evidence of significant heterogeneity for all-cause death analysis, slightly significant heterogeneity was found in the meta-analysis of recurrence and complications in our study. Regrettably, meta-regression could not explain any source, which reminded us some underlying factors may contribute to the heterogeneity. The exclusion of studies published in languages other than English was another potential limitation to our meta-analysis. It is well known that Eastern Asia and Sub-Saharan Africa were high risk areas for HCC and some articles were not published in English. Hence, these studies focused on this specific issue would not meet our inclusion criteria for our meta-analysis. These studies can have little bearing on the results of this meta-analysis, although there was no publication bias in our study.

In conclusion, this meta-analysis has demonstrated that ABT has a significantly deleterious effect on clinical outcomes of HCC patients, and is associated with an elevated risk of death, recurrence and postoperative complications. Theses findings emphasize the need for meticulous surgical technique to minimize effects of blood transfusion has been assumed to be related to the suppressive effects on the immune system. Several investigators have reported that blood transfusion suppress host immunity via a reduction in T lymphocytes function, decreased natural killer cell function, increased numbers of T-suppressor cells and decreased function of macrophages and monocytes[39;40;41]. In recent years, an increasing body of studies have demonstrated that several lymphocyte surface markers were significantly changed after ABT, and they were closely related with tumor proliferation, apoptosis and progression to metastasis. These studies have shown that the CD2 and CD4 level were decreased during postoperative period, as compared with non-transfused patients[42;43]. In addition, other studies have suggested that soluble HLA class I and soluble Fas-ligand released by leukocytes present in blood products inhibit the activity of NK cells and cytotoxic T cells, which are known to reduce immune capacity, and may predispose to postoperative infections[44;45;46].

Our study has other advantages except for the consistency of overall results. Up to now, the number of cases included in our study is one of the largest. Although publication bias is one of the major drawbacks of any meta-analysis, no publication bias was detected by Egger’s and Begg’s test, which indicated that statistical results approximated the truth. However, several limitations should be noted in our study. Firstly, all the included articles were non-randomized studies, and the evidence level is lower than that of randomized controlled trial (RCT). One reason for this is that there is no choice for patients with substantial hemorrhage but transfusion. Hence, this needs to be kept in mind when interpreting our result. In the absence of RCT now or in the future, this meta-analysis may represent a comprehensive analysis of the impact of ABT for patients with HCC. Secondly, the characteristics of patients may differ between the studies. As we know, the factors impacted the clinical outcomes of HCC patients after surgery include hepatitis C infection, pre-treatment α-fetoprotein level, liver cirrhosis and tumor stage[47;48]. Because of limited information, we were not able to control these potential confounded factors. Thirdly, surgery-related risk factors such as type and duration of operation should not be neglected. Postoperative transcathearteral chemoembolization (TACE) has been involved significantly over the past years, with significant improvement in survival[49;50]. The adjuvant treatment received by each patient was not fully understood, which could lead unknown effect to our results. Meanwhile, the lack of individual case data limited our ability to analyze the cancer-related death risk in our study. Future studies should prospectively record these data. Last but not least, although there was no evidence of significant heterogeneity for all-cause death analysis, slightly significant heterogeneity was found in the meta-analysis of recurrence and complications in our study. Regrettably, meta-regression could not explain any source, which reminded us some underlying factors may contribute to the heterogeneity. The exclusion of studies published in languages other than English was another potential limitation to our meta-analysis. It is well known that Eastern Asia and Sub-Saharan Africa were high risk areas for HCC and some articles were not published in English. Hence, these studies focused on this specific issue would not meet our inclusion criteria for our meta-analysis. These studies can have little bearing on the results of this meta-analysis, although there was no publication bias in our study.
blood loss. Therefore, to achieve a better outcome of HCC patients undergoing liver resection, ABT should not be performed if possible. In addition, the use of artificial blood products or marrow stimulants such as erythropoietin should be advocated fully. Meanwhile, the results of our study need to be interpreted in the backdrop of the subjective design of the meta-analysis with its inherent advantages and disadvantages. As a result, further randomized controlled trials would be better suited to address this issue exactly.

**Supporting Information**

**Table S1** Checklist of items to include when reporting a
this issue exactly.

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**Table S2** Newcastle-Ottawa quality assessment scale. (DOC)

**Acknowledgments**

We gratefully acknowledge the assistance of Zheng Wang from the Department of Gastroenterology, Ruijin Hospital, Shanghai Jiaotong University School of Medicine, for his assistance of literature search.

**Author Contributions**

Conceived and designed the experiments: LL, Y. Zhang. Performed the experiments: LL, ZW, SJ, BS. Analyzed the data: LL, ZW, Y. Zhang. Contributed reagents/materials/analysis tools: JL, SZ, Yilong Zhou Yuan Zhou.

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