High-Dose Chemotherapy Followed by Autologous Stem Cell Transplantation for Metastatic Rhabdomyosarcoma—A Systematic Review

Frank Peinemann1, Nicolaus Kröger2, Carmen Bartel1, Ulrich Grouven1,3, Max Pittler4, Rudolf Erttmann5, Michael Kulig1,6

1 Institute for Quality and Efficiency in Health Care (IQWiG), Cologne, Germany, 2 Interdisciplinary Clinic for Stem Cell Transplantation, University Hospital Hamburg-Eppendorf, Hamburg, Germany, 3 Hannover Medical School, Hannover, Germany, 4 German Cochrane Centre, University of Freiburg, Freiburg, Germany, 5 Department of Pediatric Hematology and Oncology, University Hospital Hamburg-Eppendorf, Hamburg, Germany, 6 Federal Joint Committee of the German Health Care System (G-BA), Berlin, Germany

Abstract

Introduction: Patients with metastatic rhabdomyosarcoma (RMS) have a poor prognosis. The aim of this systematic review is to investigate whether high-dose chemotherapy (HDCT) followed by autologous hematopoietic stem cell transplantation (HSCT) in patients with metastatic RMS has additional benefit or harm compared to standard chemotherapy.

Methods: Systematic literature searches were performed in MEDLINE, EMBASE, and The Cochrane Library. All databases were searched from inception to February 2010. PubMed was searched in June 2010 for a last update. In addition to randomized and non-randomized controlled trials, case series and case reports were included to complement results from scant data. The primary outcome was overall survival. A meta-analysis was performed using the hazard ratio as primary effect measure, which was estimated from Cox proportional hazard models or from summary statistics of Kaplan Meier product-limit estimations.

Results: A total of 40 studies with 287 transplant patients with metastatic RMS (age range 0 to 32 years) were included in the assessment. We identified 3 non-randomized controlled trials. The 3-year overall survival ranged from 22% to 53% in the transplant groups vs. 18% to 55% in the control groups. Meta-analysis on overall survival in controlled trials showed no difference between treatments. Result of meta-analysis of pooled individual survival data of case series and case reports, and results from uncontrolled studies with aggregate data were in the range of those from controlled data. The risk of bias was high in all studies due to methodological flaws.

Conclusions: HDCT followed by autologous HSCT in patients with RMS remains an experimental treatment. At present, it does not appear justifiable to use this treatment except in appropriately designed controlled trials.

Introduction

Rhabdomyosarcomas (RMS) are rare [1] malignant diseases that form a subgroup of soft tissue sarcomas (STS) which affect primarily children and young adults [2]. According to Parham 2006, "Rhabdomyosarcomas constitute a unique group of soft tissue neoplasms that share a propensity to undergo myogenesis, a well-defined biologic process that primarily occurs during embryonal and fetal development. As a result, these neoplasms tend to resemble stages of muscle formation more akin to prenatal than postnatal life" [3]. Several histologic subtypes tend to predominate in certain age groups and the embryonal and alveolar types are the most common [3]. In children 0 to 14 years of age, RMS constitute 50% of STS with an incidence rate of 4.6 per 1 million [4,5]. RMS commonly presents as a painless tumor but symptoms such as pain largely depend on the anatomical location and size of the tumor. Patients with metastatic disease at (stage 4 of TNM Classification of Malignant Tumors [1]) diagnosis have a relatively poor prognosis (5-year overall survival of 50% or lower) with current therapy such as multiagent standard-dose chemotherapy followed by surgical operation [6].

High-dose chemotherapy (HDCT) has been evaluated as an alternative treatment option for patients with metastatic RMS. The rationale for HDCT is that escalating doses of HDCT may increase survival by capturing putatively remnant malignant cells and thus overcome cell resistance to standard chemotherapy [7]. The rationale for autologous hematopoietic stem cell transplantation (HSCT) following HDCT is a planned rescue for HDCT-related severe hematologic toxicity [7]. This treatment combination has life-threatening hematologic adverse events such as graft failure,
severe infections and bleeding and non-hematologic adverse events such as multi-organ failure [8]. Of a total of 24,168 HSCT patients that were registered by the European Group for Blood and Marrow Transplantation (EBMT) in 2005, 15,276 were autologous HSCT patients and 69 were indicated for STS [9]. The potential benefit of this treatment option has not been investigated sufficiently in controlled studies [10]. Some authors have warned against the use of HDCT with autologous HSCT, indicating the possibility of repositioning of malignant cells [11]; others have questioned the rationale of HDCT with reference to the potential existence of refractory cancer stem cells [7,12,13].

The question has not been answered whether HDCT followed by autologous HSCT is able to increase overall survival in patients with RMS when compared to standard-dose chemotherapy without HSCT. Since HDCT followed by HSCT is generally not considered a treatment option for localized tumors, we concentrated on patients with metastatic RMS. The primary objective of the present systematic review is to investigate the overall survival of those patients and the secondary objective was to assess serious adverse events such as treatment-related mortality in randomized and non-randomized clinical intervention studies.

Methods

During preparation of this article we adhered to the principles and checklist of the PRISMA statement [14,15].

Study inclusion criteria

We included patients with metastatic RMS who received HDCT followed by autologous HSCT. Study design for comparative effectiveness research was limited to randomized controlled trials (RCTs) and non-randomized intervention studies [16]. We did not set a minimum number of patients (sample size) to be considered. However, we required the proportion of relevant patients to be at least 80% per study population if no stratification had been performed. If the proportion of relevant patients was less than 80% and the results were not presented separately for patients with other diseases and other interventions, then we excluded the studies. Study design for a supplemental presentation of individual data included controlled studies if aggregate data on overall survival were not reported as well as case series and case reports. Full-text publications in English language were considered. We set no limits on year of publication or year of treatment.

Search strategy

MEDLINE (1950 to 2010), EMBASE (1980 to 2010), and The Cochrane Library (to 2010) were searched without restrictions on study design, publication year, and language. The first database search was conducted 05 March 2007. An update search 05 February 2010 included a modified strategy to consider MeSH changes and exclude animal studies more precisely. We extended the search strategy by adding the term high-dose chemotherapy because the term TRANSPLANTATION was not present in title, abstract, and MeSH terms of some articles even though transplantation can be considered as their main topic. The MeSH term BONE MARROW TRANSPLANTATION was deleted from one category [17]. For 2008 MeSH there was a major revision of Publication Types (PT) and the phrase "as Topic" was added [18]. We introduced (ANIMALS not (ANIMALS and HUMANS)).sh. and replaced (ANIMALS not HUMANS).sh. To account for these changes, both searches were not restricted to publication year. The terms and the syntax used for the search in MEDLINE via Ovid as shown in Table S1 were tailored to the requirements of the other 2 databases. A final search was conducted 06 June 2010 in PubMed using MeSH Terms SARCOMA, STEM CELL TRANSPLANTATION and BONE MARROW TRANSPLANTATION, and Text Words high dose chemotherapy as search terms in various combinations. Reference lists of all included original articles and of 9 recent reviews (Banna 2007 [7]; Ek 2006 [19]; Pedrazzoli 2006 [10]; Verma 2006 [20]) were hand-searched. Information on studies registered at ClinicalTrials.gov [21], International Standard Randomised Controlled Trial Number (ISRCTN) Register [22], National Institute for Health Research UK Clinical Research Network’s (NIHR UKCRN) Portfolio Database [23], National Cancer Institute Physician Data Query (NCI PDQ) Clinical Trials [24], European Group for Blood and Marrow Transplantation (EBMT) Solid Tumor Working Party (STWP) [25] were searched online (06 June 2010) using the search terms sarcoma, stem cell, transplantation, and high dose chemotherapy in various combinations.

Study selection

To be included in the present systematic review, we required the proportion of patients with metastatic RMS as well as the proportion of patients with HDCT followed by autologous HSCT to be at least 80% per study population. In the first step, articles were excluded if the title and/or the abstract clearly referred to other diagnoses than STS and as well clearly referred to other interventions than autologous HSCT. In the second step, articles not excluded in the first step were checked whether patients with metastatic RMS were evaluated. In the third step, studies reporting about patients with metastatic RMS were categorized as controlled trials and as single-arm studies. Reporting of extractable information about overall survival was required for all included studies. For each excluded study, an appropriate reason was documented (Figure 1). All steps of the literature screening process were performed by two independent reviewers. Any disagreements were resolved by discussion.

Meta-analysis

The primary effect measure for meta-analysis of controlled studies was the hazard ratio. In case the hazard ratio was not directly given in the publication, we extracted summary statistics from Kaplan Meier product-limit estimations and estimated hazard ratios according to methods proposed by Parmar 1998 [26]. For estimation, we applied a tool, which uses p-values of the appropriate log-rank test comparing the two survival functions of interest, number of patients analyzed, and number of events in each arm [27]. If this information was not available, hazard ratios were deduced from the graphical display of the survival curves, if possible. Meta-analysis was conducted using the generic variance approach [28,29] and the random effects model [30]. Calculations were conducted using SAS version 9.2 (SAS Institute Inc., Cary, North Carolina, United States). The results of the meta-analysis were graphically displayed by means of a forest plot. Heterogeneity of the results was visually assessed and quantified using the I² value [31]. In case of considerable heterogeneity (I² > 50%) a pooled estimate is not sensible and, therefore, was not calculated [32].

To summarize outcome from individual data, Kaplan-Meier survival time estimates were extracted directly from the text or deduced from the survival curve of the publication. Individual data were pooled and an estimate for the overall survival was calculated by means of a time-to-event analysis according to the Kaplan-Meier product-limit method [33]. Calculations were conducted using the procedure Lifetest of the SAS version 9.2 (SAS Institute Inc., Cary, North Carolina, United States). The condition for including the individual data in this analysis was that information...
was provided on survival as well as at start treatment and length of follow-up for each patient.

Data collection and analysis

All steps of the data collection process were performed by two independent reviewers. Any disagreements were resolved by discussion. We collected study characteristics such as the number and region of participating centers, the treatment period, the number of analyzed patients per treatment arm; and the proportion of patients with conditions other than metastatic RMS and a regimen other than HDCT followed by autologous HSCT. Median age and gender were extracted as patients' characteristics.

The primary outcome was overall survival. Survival estimates were extracted directly from the text or deduced from the survival curve of the publication. The principal summary measure was the hazard ratio as specified in the meta-analysis section. We extracted the $p$-value of the log-rank test of the overall survival functions and the 3-year estimate of both treatment arms as well as the number of surviving patients and the median follow-up if reported.

The secondary outcomes were treatment-related mortality, secondary neoplasia, toxicity, and health-related quality of life. Treatment-related mortality was defined as deaths that have been classified by the individual publication as treatment-related deaths, or patients that have died of complications. Secondary neoplasia were considered as classified by the individual publication. HRQOL was considered if the data were measured using a questionnaire that had been validated through reporting of norms in a peer-reviewed publication. Severe adverse events grades 3 to 4 of toxicity [34] were extracted and grouped as hematological (leukopenia, neutropenia, anemia), and non-hematological (kidney, liver, nervous system, heart) toxicity.

Risk of bias

Risk of bias within studies was evaluated by assessing study design, such as retro- or prospective planning, concurrent control group, criteria for assignment of patients to treatment arms, control for confounding factors and other criteria, such as unclear selection of patients and analysis of the same patients in both treatment groups, that may increase the risk of bias especially in non-randomized trials [35]. A low risk of bias required a yes for all three of the following topics: concurrent control group, control for confounding factors, and no other risk of bias factors.

Risk of bias across studies was evaluated by assessing publication bias and outcome reporting bias. We evaluated potentially relevant studies to identify studies that may have been excluded because of missing or insufficient outcome reporting. We evaluated published study protocols to identify outcome reporting different from appropriate protocols.

Results

Search results

Of 4,784 retrieved publications, 467 full-text papers were obtained for further assessment (Figure 1). A total of 40 studies (287 transplant patients with metastatic RMS) were included. The data were structured to distinguish studies with aggregate data (controlled and non-controlled trials) from studies with individual data. Only 3 studies [36–38] (108 transplant patients, 104 controls) had a controlled, yet not randomized study design. Of the remaining 37 studies, aggregate data were extractable from 2 case series [39,40] (77 transplant patients) and individual data were extractable from 35 further case series and case reports [41–75] (102 transplant patients). We did not identify any RCTs. Two recently published studies did not meet the inclusion criteria. The number of transplanted patients and their outcome was not reported separately in a non-randomized controlled study [76]. The proportion of metastatic patients was 67% of the study population in a single-arm study [77].

Baseline data

An overview of the main study design and patients' characteristics of 3 controlled studies and of 2 non-controlled studies with aggregate data is presented in Table 1. In studies reporting aggregate data, transplantations were performed between 1991 and 2003. The studies included children and young adults ranging from 0 to 32 years of age. The number of individual data of patients with metastatic RMS included in 35 case series and case reports as well as the number of data included in a meta-analysis are shown in Table S2.

All 5 controlled studies and 1 single-arm study reported exclusively patients with metastatic RMS at diagnosis. In the other single-arm study, 1 of 7 patients did not have metastatic disease.

Primary outcome

An overview of data on overall survival is presented in Table 2. In one prospective controlled trial [38], 3-year overall survival for transplant vs. non-transplant patients was estimated at 22% vs. 55%. The difference between the treatment groups was statistically significant (log-rank $p$-value = 0.001). In a retrospective questionnaire-based controlled trial, the difference in overall survival between the treatment groups was also statistically significant but with reversed direction of effect (hazard ratio 0.38; 95% CI 0.17 to 0.88). The estimated survival after 3 years was 53% vs. 18% for transplant patients vs. non-transplant patients [37]. The estimates of another controlled trial [36] showed a numerical tendency in favor of HDCT followed by autologous HSCT (40% vs. 28%, log-rank $p$-value = 0.2), but the difference between the results of the treatment groups was statistically not significant.

We conducted a meta-analysis for overall survival of the 3 controlled trials based on hazard ratio and visualized the contradicting results in a forest plot (Figure 2). Considerable statistical heterogeneity ($I^2 = 85\%$) between the controlled studies did not justify the presentation of a pooled estimate. In Hosoi 2007 [37] patients with metastatic embryonal RMS who were younger than 10 years of age were excluded from the analysis. The number of excluded patients was not reported and constituted probably a small fraction. We did not identify statistically significantly different characteristics such as bone marrow metastasis between the treatment groups within one trial and across all 3 controlled trials.

Figure 1. Literature search and study flow. Studies: Bader 1989 [106]; Bisogno 2009 [39]; Carli 1999 [36]; Carli 2004 [107]; Dumontet 1992 [93]; Hosoi 2007 [37]; Klingebiel 2008 [38]; Koscielniak 1997 [108]; Matsubara 2003 [109]; Suitsa 2005 [110]; Yamada 2007 [40]. Abbreviations: Literature search and study flow. Abbreviations: EBMT STWP: European Group for Blood and Marrow Transplantation Soft Tissue Working Party; HDCT: high-dose chemotherapy; HSCT: hematopoietic stem cell transplantation; N: number; NCI PDQ: National Cancer Institute Physician Data Query Clinical Trials; NIH/UKRKN: National Institute for Health Research (NIHR) UK Clinical Research Network’s Portfolio Database; NRSTS: non-rhabdomyosarcoma soft tissue sarcoma; RMS: rhabdomyosarcoma; STS: soft tissue sarcoma. doi:10.1371/journal.pone.0017127.g001
Overall survival estimates 3 years after HDCT followed by autologous HSCT reported in 2 single-arm studies with aggregate data were at 28% [40] and at 42% [39]. These estimates were within the range for this treatment option (22% to 53%), which was reported in the 3 controlled trials. We conducted a meta-analysis of the pooled individual outcome data (Table 2). The Kaplan-Meier product-limit estimation resulted in a pooled estimate 3 years after treatment at 29% (95% CI: 15 to 43) (Figure 3). Again, this value was within the range of the 3 controlled trials. It should be noted that only 45% (45 out of 101 patients) of all included individual data were suitable for this analysis.

Secondary outcomes

Extractable data on treatment-related mortality were reported for 10 transplant patients in 8 studies (Table 3) and secondary neoplasia for 1 transplant patient in 1 study (Table 3). There was a lack of extractable data on severe (grade 3 to 4) non-hematological toxicity in a large proportion of the included studies. Two patients with kidney and 1 patient with heart toxicity were reported. In 3 studies, liver toxicity was not present in all evaluated patients (Table 3). Due to the low quantity of analyzable data, it was not possible to reasonably evaluate toxicity. A study on health-related quality of life [78] could not be included in the benefit assessment because the results were not connected to each individual diagnosis but rather reported for a group of transplant-ed patients with various diseases. We did not identify any studies with health-related quality of life endpoints.

Risk of bias

Risk of bias within studies was high for all included studies mainly due to flaws of study design, assignment of patients to treatment groups, and missing control for confounding (Table 4). Due to historical controls in the study design used in the study of Carli 1999 [36], patients in the control group were recruited earlier in time than patients who received HDCT followed by autologous HSCT. This allocation to treatment groups by time differences may introduce selection bias [79]. Only data of 96

Table 1. Study and patient’s characteristics.

| Study                  | N. centers/country | Treatment period; years | N. evaluable patients* | Metastatic RMS; % | Age; median years (range) | Gender; % males |
|------------------------|--------------------|-------------------------|------------------------|-------------------|--------------------------|-----------------|
| Non-randomized controlled studies |                    |                         |                        |                   |                          |                 |
| Carli 1999/2004 [36,107] | 5/Europe           | HSCT 1991 to 1995 vs. (historical control 1989 to 1991) | 52 vs. 44 | 100 vs. 100 | (0 to 18) | – |
| Hosoi 2007 [37]         | 63/Japan           | 1991 to 2002            | 22 vs. 20 | 100 vs. 100 | (0 to 20) | – |
| Klingebiel 2008 [38]    | 50/2/Europe        | 1995 to 2003            | 34 vs. 40 | 100 vs. 100 | (under 22) | – |
| Single-arm studies with aggregate data                  |                    |                         |                        |                   |                          |                 |
| Bisogno 2009 [39]       | 1/Italy            | –                       | 70^2                 | 100              | (0 to 20) | 47 |
| Yamada 2007 [40]        | 1/Japan            | –                       | 7                     | 86 (6/7)         | (15 to 32) | 29 |

^information not reported in the publication or not applicable.
*for non-randomized controlled studies: HSCT vs. control.
^Bisogno 2009: 11% patients (8 of 70 patients) without HSCT included.
Abbreviation. HSCT: hematopoietic stem cell transplantation; RMS: rhabdomyosarcoma.
doi:10.1371/journal.pone.0017127.t001

Table 2. Overall survival.

| Study                  | Follow-up begins at | 3-year overall survival (95% CI if reported), HSCT vs. control; % | Statistically significant difference |
|------------------------|---------------------|------------------------------------------------------------------|-------------------------------------|
| Non-randomized controlled studies |                    |                                                                  |                                     |
| Carli 1999/Carli 2004^ [36,107] | diagnosis          | 40 (26 to 55) vs. 28 (13 to 42)\(^\dagger\) | no (p = 0.200)^2                     |
| Hosoi 2007 [37]        | diagnosis           | 53 vs. 18^3                                                      | YES (0.38 (0.17 to 0.88))^\|
| Klingebiel 2008 [38]   | diagnosis           | 22 vs. 55^4                                                      | YES (p = 0.001)^3                    |
| Single-arm studies with aggregate data                  |                    |                                                                  |                                     |
| Bisogno 2009 [39]       | diagnosis           | 42 (31 to 54)                                                   | not applicable                       |
| Yamada 2007 [40]        | treatment           | 28                                                               | not applicable                       |
| Pooled individual data                  |                    |                                                                  |                                     |
| 45 patients (19 studies) | treatment           | 29 (15 to 43)                                                   | not applicable                       |

^Carli 2004 follow-up paper of Carli 1999.
^Carli 2004: 5-year overall survival for HSCT 36% (95% CI 23% to 49%) vs. control 27% (95% CI 14% to 41%).
^p-value of log-rank test.
^Hosoi 2007 and Klingebiel 2008: 3-year survival estimates deduced from the graphical display of the survival curves.
^Hazard ratio (95% confidence interval).
Abbreviation. CI: confidence interval; HSCT: hematopoietic stem cell transplantation; RMS: rhabdomyosarcoma.
doi:10.1371/journal.pone.0017127.t002

To conclude, HDCT followed by autologous HSCT provided promising survival outcomes with within-range estimates of 28% to 42%. However, the risk of bias was high, primarily due to flaws in study design, assignment, and confounding factors. Further studies with higher-quality designs are needed to accurately evaluate the benefits and risks associated with this treatment option. Additionally, the lack of extractable data on severe non-hematological toxicities limits the comprehensive assessment of this therapy. Future research should focus on improving study methodology and ensuring adequate data collection to better understand the efficacy and safety of HDCT followed by autologous HSCT in the management of RMS.
Mortality (HSCT vs. CT)
Effect: hazard ratio
Random effects model - DerSimonian and Laird (for presentation of the weights)

| Study          | Logarithmic Effect | SE   | Effect (95% CI) | Weight | Effect | 95% CI |
|----------------|-------------------|------|-----------------|--------|--------|--------|
| Carlsson 1999 / Carlsson 2004 | 0.00              | 0.28 | 35.3            | 1.00   |        | [0.60, 1.87] |
| Hosoi 2007     | -0.97             | 0.42 | 30.1            | 0.38   |        | [0.17, 0.86] |
| Klingebiel 2008| 0.64              | 0.29 | 34.0            | 2.02   |        | [1.32, 4.93] |

Heterogeneity: Q=13.26, df=2, p=0.001, I²=84.9%

Figure 2. Meta-analysis of overall survival of 3 controlled studies. A meta-analysis of overall survival of 3 controlled trials based on hazard ratio was conducted. The forest plot shows conflicting results between these trials. Considerable heterogeneity did not justify a pooled estimate.

doi:10.1371/journal.pone.0017127.g002

(55%) of 175 enrolled metastasizing patients were analyzed because complete remission before the third chemotherapy cycle was required for inclusion. Consequently, data of 78 patients were not considered because they did not achieve the required complete remission. Data of one patient were excluded because therapy information was missing.

Klingebiel 2008 [38] analyzed 96 (33%) of 295 enrolled metastasizing patients with soft tissue sarcoma in a prospective multicenter trial. Data of 37 patients were not included because of missing treatment information, data of 162 patients were not included because of incomplete treatment. The proportion of RMS patients among 199 enrolled but not analyzed patients was not reported. Of 96 analyzed patients, 22 patients had a diagnosis other than RMS and the rest of 74 patients with metastatic RMS were investigated in a subgroup analysis. Patients in the control group were treated with an oral maintenance treatment, which may not be a typical example of conventional chemotherapy and the outcome may not be comparable with those of other studies. In contrast to the overall study result, in a subgroup analysis of 43 RMS patients with bone and/or bone marrow involvement the overall survival was not different between test and control group.

Hosoi 2007 [37] defined a high-risk group of 42 (13%) metastasizing RMS patients (TNM IV) of a total of 331 enrolled patients with RMS. These data were collected from 63 institutions after a questionnaire had been sent to 93 institutions. A non-responder analysis was not reported. Selection bias may be considerable due to a questionnaire-based analysis. Patients with metastasizing embryonal RMS who were younger than 10 years of age were analyzed in another risk group. Exclusion of these patients from the high-risk group may have influenced the outcome and reduced its comparability with those of other studies.

Figure 3. Meta-analysis of pooled individual data of single arm studies. Kaplan-Meier product-limit estimation of overall survival was conducted using individual data of patients with metastatic RMS (total 45, failed 31, censored 14) from 19 case series and case reports. Information about outcome (dead or alive) and follow-up (time of survival after begin of treatment) was required for each individual. A considerable proportion of 56% (57/102 patients) of all included individual data was not considered because appropriate information was not extractable. Number of subjects at risk after each additional year of follow-up.

doi:10.1371/journal.pone.0017127.g003
Risk of bias across studies was difficult to evaluate because the number of trials included in the meta-analysis was too small for conducting a funnel plot. We did not identify significant outcome reporting bias.

**Discussion**

**Primary outcome**

We conducted a meta-analysis based on hazard ratios of overall survival and found conflicting results in 3 controlled trials. A pooled estimate was not presented because statistical heterogeneity can be categorized as considerable [80]. Heterogeneity was not explained by the distribution of clinical characteristics between treatment groups. Outcome data varied considerably between studies and the 3-year overall survival estimates after HDCT followed by autologous HSCT ranged from 22% to 52% in the 3 controlled trials. Results from single-arm studies based on aggregate data or on pooled individual data were within this range.

To our knowledge, this is the first systematic review and meta-analysis that compares HDCT followed by autologous HSCT vs. standard-dose chemotherapy without transplantation in patients...
with metastatic RMS. Weigel 2001 [81] reported results from single-arm studies on HDCT followed by autologous HSCT in patients with metastatic and relapsed RMS. Verma 2008 [20] investigated HDCT followed by autologous HSCT in patients with inoperable, locally advanced, or metastatic STS, and did not address specifically on RMS. Admiraal 2007 [82] conducted a systematic review exclusively on patients with metastatic RMS. Secondary outcome

Secondary outcome

Treatment-related mortality was extractable from only few studies. Early (<100 days) and late mortality (>100 days) was not differentiated in the studies. Secondary neoplasia was extractable from only a few studies. As events are expected to cumulate with time, the incidence rate of secondary neoplasia in the present systematic review may be underestimated because of a short follow-up time and of loss of patients to follow-up. Treatment-related mortality, secondary neoplasia, and toxicity proportion was not summarized due to scarcity of data.

Outcome reporting bias

Outcome reporting bias [83] is defined as the selection of a subset of the original recorded outcome variables for publication. Systematic reviews require addressing the issue of missing outcome data because outcome reporting bias can affect their conclusions [84]. We did not identify significant outcome reporting bias. All 3 controlled trials reported statistical summary data such as hazard ratio and \( p \)-values of log-rank test. Incomplete data as well as a short follow-up may have compromised the results.

Study publication bias

Study publication bias is defined as publication of research results depending on their results [85]. The number of trials included in the meta-analysis was too small for conducting a funnel plot analysis because at least 10 trials are required for a reasonable analysis [86]. The strengths of the present systematic review are the broadness of the search strategy and the comprehensiveness of the published data included. Nevertheless, there may be a slight possibility that an unknown number of studies were not registered and not published.

Language bias

Results in English language articles could be different from those of articles written in another language [87]. Non-English language requires expensive translation to prevent selective outcome extraction and misinterpretation of results. Funding for translation was not provided and we excluded all non-English including German language articles. Restricting the inclusion of studies to English language may have little effect on summary treatment effect estimates [88,89] and German language articles may not play a preeminent role in the dissemination of medical research [90].

Internal validity

We identified a high risk of bias within all studies. The criteria of allocation of the patients to the treatment groups were not reported in all included studies. Klingebiel 2008 [38] reported the outcome of interest in a subgroup analysis but did not present baseline data for this subgroup. Hosoi 2007 [37] and Carli 1999 [36] did not report median age and distribution of gender. The contradictory results reported by Klingebiel 2000 [30] vs. Hosoi 2007 [37] call for an explanation. Both authors reported patients with metastatic RMS in a subgroup. Hosoi 2007 [37] excluded patients with metastatic embryonal RMS younger than 10 years of age but Klingebiel 2000 [38] did not. Baseline data for the subgroup were not sufficient to specify the number of excluded patients. The 2 studies [Klingebiel 2008 [38] vs. Hosoi 2007 [37]] had a different setting (Germany vs. Japan), a different design (prospective vs. retrospective), and a different data collection process (protocol driven clinical trial vs. questionnaire). We did not find any statistically significant differences of the patients’ characteristics between the treatment groups. Given the data presented, it remains a matter of speculation whether the risk of bias within the studies might explain the different outcome as we do not know in what direction the results were possibly biased.

Heterogeneity

Strict adherence to inclusion criteria limited clinical heterogeneity to a low to moderate range. Patient’s characteristics such as clinical stage, age, and histological subtypes were roughly comparable between the treatment groups.

Lack of controlled trials

For sarcomas, at least 20 articles [8,11,19,54,81,91–105] addressed the need for conducting RCTs to evaluate the role of HDCT followed by autologous HSCT vs. standard-dose chemotherapy. Consequently, the lack of RCTs and the lack of clinical controlled trials addressing the issue is surprising indeed. Only 3 non-randomized controlled studies of poor quality were published to date and are included in the present review. According to the European Group for Blood and Marrow Transplantation (EBMT), 69 autologous HSCT were assigned to patients with soft tissue sarcoma in the year 2005 [9].

Conclusions

High-dose chemotherapy followed by autologous hematopoietic stem cell transplantation in patients with metastatic rhabdomyosarcoma remains an experimental treatment. At present, it does not appear justifiable to use this treatment except in appropriately designed controlled trials.

Supporting Information

Table S1 Search strategy used in MEDLINE via Ovid. (DOCX)

Table S2 Studies with individual data. (DOCX)

Author Contributions

Conceived and designed the experiments: FP. Performed the experiments: FP CB. Analyzed the data: FP CB MK. Contributed reagents/materials/analysis tools: UG. Wrote the paper: FP NK CB UG MP RE MK.

References

1. Sobin LH, Gospodarowicz MK, Wittekind C (2009) UICC: TNM classification of malignant tumors. Oxford: Wiley-Blackwell.
2. Weiss SW, Goldblum JR (2001) Enzinger and Weiss’s soft tissue tumors. Saint Louis: Mosby.
81. Weigel BJ, Breitfeld PP, Hawkins D, Crist WM, Baker KS (2001) Role of high-dose chemotherapy in advanced pediatric solid tumors. Pediatr Hematol Oncol 18: 121–132.

80. (2009) Section 9.5.2 Identifying and measuring heterogeneity. In: Higgins JPT, Green S, eds. Cochrane Handbook for Systematic Reviews of Interventions Version 5.0.2 [updated September 2009]. The Cochrane Collaboration.

79. Stiff PJ, Agovi MA, Antman KH, Blaise D, Camitta BM, et al. (2010) High-dose chemotherapy with blood or bone marrow rescue in the treatment of metastatic or recurrent rhabdomyosarcoma. J Pediatr Hematol Oncol 32: 272–276.

78. Alm J, Breitfeld PP, Hawkins D, Crist WM, Baker KS (2001) Role of high-dose chemotherapy with autologous hematopoietic stem cell transplantation in the treatment of metastatic or recurrent rhabdomyosarcoma. Pediatr Hematol Oncol 18: 121–132.

77. Stiff PJ, Agovi MA, Antman KH, Blaise D, Camitta BM, et al. (2010) High-dose chemotherapy with blood or bone marrow rescue in the treatment of metastatic or recurrent rhabdomyosarcoma. J Pediatr Hematol Oncol 32: 272–276.

76. Alm J, Breitfeld PP, Hawkins D, Crist WM, Baker KS (2001) Role of high-dose chemotherapy with autologous hematopoietic stem cell transplantation in the treatment of metastatic or recurrent rhabdomyosarcoma. Pediatr Hematol Oncol 18: 121–132.

75. Kavan P, Stankova J, Koutecky J, Gajdos P (1997) High-dose chemotherapy and autologous peripheral blood stem cell transplantation for children with solid tumors. Bone Marrow Transplant 21: 526–530.

74. Walterhouse DO, Hoover ML, Marymont MA, Kletzel M (1999) High-dose chemotherapy with autologous hematopoietic stem cell transplantation in children with solid tumors. Bone Marrow Transplant 22: 355–360.

73. Ritchie DS, Grigg AP, Roberts AW, Rosenthal MA, Fox RM, et al. (2004) High dose chemotherapy followed by autologous peripheral blood stem cell rescue for children with solid tumours. Bone Marrow Transplant 33: 1097–1105.

72. Sanz N, de Mingo L, Florez F, Rollan V (1997) Rhabdomyosarcoma of the brain: a retrospective review of patients treated at the hospital for sick children between 1989 and 1999. J Pediatr Hematol Oncol 21: 1043–1047.

71. Sakayama K, Tauchi H, Sugawara Y, Kidani T, Tokuda K, et al. (2008) A long-term follow-up of 103 children with stage IV rhabdomyosarcoma: the experience at Chicago Children’s Memorial Hospital. Pediatr Blood Cancer 50: 783–784.

70. Boughton ME, Gurney JG, Mertens AC, Smith MA, Stovall M, et al. (2008) Rhabdomyosarcoma: where do we go from here. J Pediatr Hematol Oncol 23: 266–267.

69. Hale GA (2005) Autologous hematopoietic stem cell transplantation for pediatric solid tumors. Expert Rev Anticancer Ther 5: 835–846.

68. Ritchie DS, Grigg AP, Roberts AW, Rosenthal MA, Fox RM, et al. (2004) High dose chemotherapy followed by autologous peripheral blood stem cell rescue for children with solid tumours. Bone Marrow Transplant 33: 1097–1105.

67. Bader JL, Horowitz ME, Dewan R, Watkins E, Triche TJ, et al. (1989) Phase II study of high-dose chemotherapy and hematopoietic stem cell transplantation in children with solid tumors. J Pediatr Surg 38: 130–133.

66. Mitchell PL, Shepherd VR, Proctor HM, Dainton M, Cabral SD, et al. (1994) Peripheral blood stem cells used to augment autologous bone marrow transplantation. Arch Dis Child 70: 237–240.

65. Munoz LL, Wharam M, Kaiser H, Leventhal BG, Ruyffmann F (1983) Magna-Flux irradiation and autologous marrow rescue in the treatment of pediatric solid tumors. Int J Radiat Oncol Biol Phys 9: 1951–1954.

64. Navid F, Santana VM, Billups CA, Merchant TE, Furman WL, et al. (2006) Concomitant administration of vincristine, doxorubicin, cyclophosphamide, dacarbazine, and prednisone for high-risk rhabdomyosarcoma. Cancer 106: 1046–1056.

63. Mitchell PL, Shepherd VR, Proctor HM, Dainton M, Cabral SD, et al. (1994) Peripheral blood stem cells used to augment autologous bone marrow transplantation. Arch Dis Child 70: 237–240.

62. Sano N, de Mingo L, Flores R, Rollan V (1997) Rhabdomyosarcoma of the biliary tree. Pediatr Surg Int 12: 290–291.
patients with high-risk rhabdomyosarcoma who achieve complete remission with conventional chemotherapy. Pediatr Hematol Oncol 20: 201–210.

110. Suita S, Noguchi S, Takamatsu H, Mizote H, Nagasaki A, et al. (2005) Clinical characteristics and the prognosis of rhabdomyosarcoma - A report from the Study Group for Pediatric Solid Malignant Tumors in the Kyushu Area, Japan. Eur Journal Pediatr Surg 15: 409–413.