A theoretician’s analysis of the supernova data and the limitations in determining the nature of dark energy

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ABSTRACT
Current cosmological observations show a strong signature of the existence of a dark energy component with negative pressure. The most obvious candidate for this dark energy is the cosmological constant (with the equation of state \( w = p/\rho = -1 \)), which, however, raises several theoretical difficulties. This has led to models for dark energy component which evolves with time. We discuss certain questions related to the determination of the nature of dark energy component from observations of high redshift supernova. The main results of our analysis are: (i) Even if the precise value of \( w_X \) is known from observations, it is not possible to determine the nature of the unknown dark energy source using only kinematical and geometrical measurements. We have given explicit examples to show that different types of sources can give rise to a given \( w_X \). (ii) Although the full data set of supernova observations (which are currently available) strongly rule out models without dark energy, the high (\( z > 0.25 \)) and low (\( z < 0.25 \)) redshift data sets, individually, admit decelerating models with zero dark energy. Any possible evolution in the absolute magnitude of the supernovae, if detected, might allow the decelerating models to be consistent with the data. (iii) We have introduced two parameters, which can be obtained entirely from theory, to study the sensitivity of the luminosity distance on \( w_X \). Using these two parameters, we have argued that although one can determine the present value of \( w_X \) accurately from the data, one cannot constrain the evolution of \( w_X \).

1 INTRODUCTION
One of the equations governing the dynamics of a Friedman universe, \((\ddot{a}/a) = -(4\pi G/3)(\rho + 3p)\), implies that the universe will accelerate \((\ddot{a} > 0)\) if \((\rho + 3p) < 0\). The analysis of high-redshift supernova data (Riess et al. 1998; Perlmutter et al. 1999; Riess 2000) seems to suggest that \(\ddot{a}\) is indeed positive thereby requiring \((\rho + 3p) < 0\). This condition requires at least \(\rho\) or \(p\) to be negative. Cosmologists have not yet become desperate enough to suggest \(\rho < 0\); but there was remarkably low resistance in the community to accepting the existence of a constituent with \( w \equiv (p/\rho) < -(1/3) \). This was, to a great extent, facilitated by the fact that cosmologists have long since toyed with the existence of a non-zero cosmological constant with the stress-tensor \( T_k^k = \rho_X \delta_k^k \) corresponding to \( w_X = -1 \). In fact, much before the arrival of supernova data on the scene, there were indications that the universe actually may be populated by a energy density which is very smoothly distributed over large scales. A very clear argument to this effect was given based on the analysis of APM data in 1990 (Efstathiou, Sutherland, & Maddox 1990). Later, the first analysis of the COBE data in 1992 indicated that the shape of the standard cold dark matter (SCDM) spectrum needs to be modified and the existence of a smoother energy distribution, like the cosmological constant will be required (Padmanabhan & Narasimha 1992; Efstathiou, Bond, & White 1992). An analysis of a host of observations available in 1996 was used to suggest that the data supports the existence of a non-zero cosmological constant (Bagla, Padmanabhan, & Narlikar 1996). Given this background, it was not surprising that the cosmologists were not too shocked by the supernova data which became available from 1998 onwards. Currently, there are also other observations, like those of the age of the universe, gravitational lensing surveys etc., which strongly suggest the presence of a positive cosmological constant [see Padmanabhan (2002a) for a comprehensive review and references].

There are, however, well known deep theoretical problems with the existence of a non-zero cosmological constant \( \Lambda \) with a magnitude of about \( \Lambda (G\hbar/c^2) \approx 10^{-123} \). These are well documented in the literature and will not be discussed here; for a recent review see Padmanabhan (2002a). This has prompted a host of activity in which one looks for a dark energy component in the universe with \( w < -(1/3) \) which is different from cosmological constant [see Padmanabhan (2002a) and references cited therein]. If \( w_X(a) \) denotes the time dependent equation of state parameter for an unknown dark energy component in the universe, then the equation of motion \( d(\rho_X a^3) = w_X \rho_X a^3 \) integrates to give
\[
\rho_X(a) = \rho_X(a_0) \exp \left\{ -3 \int_{a_0}^a \frac{da'}{a'} [1 + w_X(a')] \right\}
\]
If \( w_X \) is not identically equal to \(-1\), then the dark energy density will evolve with time (even if \( w_X \) is a constant). This suggests the possibility that the cosmological “constant” can be time dependent thereby alleviating some of the difficulties which arises in the case of \( w_X = -1 \). Several such models (usually based on scalar fields) have been constructed in the literature, all of which, generally have a time dependent \( w_X(a) \). The introduction of such an ad-hoc, time dependent function into cosmology, of course, takes
away much of the credibility; nevertheless, this issue needs to be settled ultimately by observations and not by aesthetics. (If the observations forces us into a solution which is unaesthetic by current standards, we will either conveniently change the standards or live with it!)

In this paper, we discuss certain questions related to the determination of the nature of dark energy component from observations related to supernova. This issue has been studied in great detail by several groups and our contribution in this paper will border on pedagogy and will present a particular point of view. In the next section we briefly recall and stress some inherent theoretical degeneracies in these analysis. In Section 3 we reanalyze the currently available supernova data in order to focus attention on some key elements. (These results exist in alternate forms in published literature but we believe our analysis brings out some features clearly.) Based on the lessons learnt in this section, we carry out a similar analysis for possible future supernova observations and point out what can and cannot be achieved. We have intentionally kept the data analysis at a fairly simple level (indicated in the title by ‘a theoretician’s analysis’). We subscribe to the point of view that any result which cannot be revealed by a simple analysis of data, but arises through a more complex statistical procedure, is inherently suspect and a conclusion as important as the existence of dark energy with a more complex statistical procedure, is inherently suspect and a serious — degeneracy which seems to have been inadequately stressed in the literature. Let us assume that the universe is made of two components: \( \rho_{\text{dm}}(a) \), which is known from independent observations and a component \( \rho_X(a) \) which is not known. From the Friedmann equation, it follows that

\[
\frac{8\pi G}{3} \rho_X(a) = H^2(a) [1 - \Omega_{\text{dm}}(a)]; \quad \Omega_{\text{dm}}(a) \equiv \frac{8\pi G \rho_{\text{dm}}(a)}{3H^2(a)}
\]

Taking a derivative of \( \ln \rho_X(a) \) and using (1), it is easy to obtain the relation

\[
\frac{1}{3} \frac{d}{d \ln a} \ln \left[ \left( 1 - \Omega_{\text{dm}}(a) \right) H^2(a) a^3 \right] = w_X(a)
\]

If geometrical observations of the universe give us \( H(a) \) and other observations give us \( \rho_{\text{dm}}(a) \) then one can determine \( \Omega_{\text{dm}}(a) \) and thus \( w_X(a) \). This is the best uniform background cosmology can do for us. As far as cosmological model building is concerned, one is usually interested in the features of the dark energy “fluid”, and it is usually sufficient to know \( w_X(a) \) and the speed of sound of the fluid. However, theoretical physicists would like to know something more about the nature of the dark energy “fluid” — in particular, they would find it useful to determine the Lagrangian for the fluid. (For example, in the case of inflation, a considerable amount of effort is being spent to understand the form of the potential for the scalar field.) In this paper, we shall take the viewpoint of a theoretical physicist, where the nature of the dark energy is defined by its Lagrangian.

Now, if a geometrical observation in future suggest that \( w_X(a) \) is not identically equal to \(-1\), then the question arises as to what is the nature of this dark energy field. It seems virtually impossible that its nature can be determined from laboratory experiments (unlike, for example, WIMPS which might constitute dark matter). However, knowing \( w_X(a) \) is grossly inadequate for determining the physical nature of dark energy. For example, even if one makes another gigantic leap of faith and assumes that the dark energy arises from a scalar field, it is possible to come up with scalar field Lagrangians of different forms leading to same \( w_X(a) \) [see Padmanabhan (2002b)]. To illustrate this point, we shall discuss two possibilities:

\[
L_{\text{quint}} = \frac{1}{2} \partial_i \phi \partial^i \phi - V(\phi), \quad L_{\text{tach}} = -V(\phi) \sqrt{1 - \partial_i \phi \partial^i \phi}
\]
Both these Lagrangians involve one arbitrary function $V(\phi)$. The first one, $L_{\text{quin}}$, which is a natural generalization of the Lagrangian for a non-relativistic particle $[L = q^2/2 - V(\phi)]$, is usually called quintessence. Similarly, the second one, $L_{\text{tach}}$, is a generalization of the Lagrangian for a relativistic particle $[L = -m\sqrt{1 - \dot{q}^2}]$, and is usually called a tachyonic Lagrangian. It arises naturally in the string theoretical context; see Sen (2002a); Sen (2002b); Sen (2002c). For a discussion of the cosmological aspect of this Lagrangian, see Padmanabhan & Choudhury (2002) and references therein. When these Lagrangians act as sources in Friedmann universe, they are characterized by density and equation of state parameters given by

$$\rho_{\text{quin}, a} = \frac{1}{2} \dot{\phi}^2 + V, \quad w_{\text{quin}, a} = \frac{1 - (2V/\dot{\phi}^2)}{1 + (2V/\dot{\phi}^2)}$$

(8)

and

$$\rho_{\text{tach}, a} = \sqrt{1 - \dot{\phi}^2}, \quad w_{\text{tach}, a} = \frac{\dot{\phi}^2 - 1}{\sqrt{1 - \dot{\phi}^2}},$$

(9)

respectively.

Since both the Lagrangians have one undetermined function $V(\phi)$, it is possible to choose this function in order to produce a given $w_X(a)$ [or equivalently, $H(a)$]. For a flat universe, one can determine the function $V(\phi)$ for the quintessence model from the implicit relations (Ellis & Madsen 1991)

$$V(a) = \frac{H(a)}{16\pi G}[1 - \Omega_{ka}(a)]$$

$$\times \left[\frac{H(a) + 2aH'(a) - aH(a)\Omega_{ka}(a)}{1 - \Omega_{ka}(a)}\right]^{1/2}$$

(10)

$$\phi(a) = \sqrt{\frac{1}{8\pi G}} \int \frac{da}{a}$$

$$\times \left[ a\Omega_{ka}(a) - [1 - \Omega_{ka}(a)] \frac{d\ln H^2(a)}{d\ln a} \right]^{1/2}$$

(11)

Similarly, in the case of tachyonic scalar field, the potential function is given by (Padmanabhan 2002b)

$$V(a) = \frac{3H^2(a)}{8\pi G} [1 - \Omega_{ka}(a)]$$

$$\times \left\{ 1 + \frac{2aH'(a)}{3H(a)} - \frac{a\Omega_{ka}(a)}{3[1 - \Omega_{ka}(a)]} \right\}^{1/2}$$

(12)

$$\phi(a) = \int \frac{da}{aH(a)}$$

$$\times \left\{ \frac{a\Omega_{ka}(a)}{3[1 - \Omega_{ka}(a)]} - \frac{2aH'(a)}{3H(a)} \right\}^{1/2}$$

(13)

We can present an explicit example of a situation where $w_X$ is assumed to be known of the simple form

$$w_X(z) = w_0 - w_1(a - 1) = w_0 + w_1 \frac{z}{1 + z},$$

(14)

where $w_0$ measures the current value of the parameter and $-w_1$ gives its rate of change (with respect to the scale factor) at the present epoch. In addition to simplicity, this parametrization has the advantage of giving finite $w_X$ in the entire range $0 < z < \infty$. We shall use this simple parametrization later to discuss the possibility of constraining $w_X$ from observations. The Hubble parameter in this case is given by

$$\frac{H^2(z)}{H_0^2} = \frac{\Omega_m(1 + z)^3}{(1 - \Omega_m)(1 + z)^3(1 + w_0 + w_1) e^{-3w_1 z/(1 + z)}}.$$
the quintessence and tachyonic fields. The corresponding potentials $V(\phi)$ can be determined from equations (11) and (12). Unfortunately, the relations cannot be written in a closed form (unless $w_1 = 0$) in this case. The form of the potential for both the cases can be calculated numerically for given values of $w_0$ and $w_1$, which are plotted in Figure 1. The left frame shows $V(\phi)$ for the quintessence field while the right frame shows that for the and tachyonic field. The potentials are plotted for the parameter values $\Omega_m = 0.3, \Omega_X = 1 - \Omega_m = 0.7, w_0 = 0.5, w_1 = -0.1$ using equations (10) - (13). Since the equations give the potential $V(\phi)$ in the parametric form $[V(z), \phi(z)]$, each point on the $V - \phi$ curve can be labelled by the corresponding value of redshift, $z$. We have pointed some particular redshifts in both the curves. It is clear that the scalar field starts from a large value of the potential, rolls down as time progresses and the density contributed by the potential reaches a value $\sim \Omega_X$ at $z = 0$.

This discussion shows that even when $w_X(a)$ or $H(a)$ is known, it is not possible to proceed further and determine the nature of the unknown dark energy source using only kinematical and geometrical measurements. This constitutes a serious theoretical problem in cosmology if observations suggest that $w_X$ is not identically equal to $-1$ (Huterer & Turner 1999).

3 WARM UP: CURRENT supernova DATA AND THEIR ANALYSIS

We shall next reanalyze the currently available supernova data in order to stress certain features which are inherent in such an analysis. This may be considered a warm up exercise for the next section.

The observations directly measure the apparent magnitude $m$ of a supernova and its redshift $z$. The apparent magnitude $m$ is related to the luminosity distance $d_L$ of the supernova through

$$m(z) = M + 5 \log_{10} \left( \frac{d_L(z)}{1 \text{ Mpc}} \right) + 25,$$

where $M$ is the absolute magnitude (which is believed to be constant for all supernovae of Type-Ia — this is what is called the “standard candle hypothesis”). It is convenient to work with a dimensionless quantity (called the “Hubble-constant-free” luminosity distance)

$$Q(z) \equiv \frac{H_0}{c} d_L(z)$$

which gives

$$m(z) = M + 5 \log_{10} Q(z),$$

where

$$M = M + 5 \log_{10} \left( \frac{c/H_0}{1 \text{ Mpc}} \right) + 25 = M - 5 \log_{10} h + 42.38.$$ (19)

Any model of cosmology will predict $Q(z)$ with some undetermined parameters (say, for example, $\Omega_m, \Omega_\Lambda$). These parameters, along with $M$ (which is itself a combination of the absolute magnitude of the Type-Ia supernova and the Hubble constant), are obtained by comparing with observations. According to the standard hot big bang model of cosmology [see e.g., Padmanabhan (2002c)],

$$Q(z) = (1 + z) \frac{\sinh[z(1 + z)]}{\sqrt{\Omega_k}},$$

where

$$x(z) = \int_0^z \frac{dz'}{H(z')},$$

The data in this paper is based on the redshift-magnitude relation of 54 Type-Ia supernovae [excluding 6 outliers from the full sample of 60, which consists of 18 low-redshift points from the Calán/Tololo supernova survey (Hamuy et al. 1996) and 42 points from the Supernova Cosmology Project (Perlmutter et al. 1999)] and that of supernova 1997ff at $z = 1.755$ (Riess et al. 2001). The values of $m$ used have been already corrected for the supernova light curve width-luminosity relation, galactic extinction and possible K-correction. In addition, the magnitude for supernova 1997ff has been corrected for lensing effects (Benitez et al. 2002). Given a cosmological model with some free parameters, we have obtained the best-fit parameter values and the corresponding covariance matrix using the Levenberg-Marquardt method (Press et al. 1992). The details of this fitting procedure, although quite well-known, are given in Appendix A for completeness.

Let us start with flat models with $\Omega_m + \Omega_\Lambda = 1; \Omega_k = 0$, which are currently favoured strongly by CMBR data (for recent results, see Sievers et al. 2002). The simple analysis mentioned above gives a best-fit value of $\Omega_m$ (after marginalizing over $M$) to be $0.31 \pm 0.08$. The best-fit $M$ (after marginalizing over $\Omega_m$) is found to be $23.95 \pm 0.05$ (all the errors quoted in this paper are 1$\sigma$). The comparison between these flat models and the observational data is shown in Figure 2. In order to determine $Q(z)$ for the observational data points, we have used equation (18) with the best-fit value $M = 23.95$.

Although the best-fit analysis shows that the data favour strongly for a positive non-zero $\Omega_\Lambda$ which, in turn, implies the presence of an accelerating universe, the same conclusion is not visually obvious from Figure 2. In order to see that the data favours models with non-zero $\Omega_\Lambda$, one usually plots the supernova magnitude with respect to a fiducial best-fit model [see, for example, Perlmutter et al. (1999) and Riess et al. (2001)].
In particular, the decelerating model ($\Omega_m = 1$) is quite consistent with both the low and high redshift data sets when they are treated separately. One needs to combine the low and high redshift data to constrain $\Omega_m$ – because of the angular slant of the ellipses, a best-fit region around $\Omega_m = 0.31$ is isolated, as can be seen from the left panel. [This analysis indirectly stresses the importance of any evolutionary effects. If, for example, supernova at $z \gtrsim 0.25$ and supernova at $z \lesssim 0.25$ have different absolute luminosities because of some unknown effect, then the entire data set can be made consistent with $\Omega_m = 1$, $\Omega \Lambda = 0$ model. It certainly appears ad-hoc; but one should compare the ad-hocness in introducing a dimensionless constant $\Lambda(Gh/c^4) \approx 10^{-123}$ in the physical system to explain the cosmological observations (Padmanabhan 2002a)]

We now generalize some of the results for non-flat cosmologies. In this case, we have three free parameters, namely, $\Omega_m$, $\Omega \Lambda$ and $\mathcal{M}$. The confidence region ellipses in the $\Omega_m - \Omega \Lambda$ plane can be drawn in two ways – (i) by keeping $\mathcal{M}$ as a free parameter and then marginalize the joint probability distribution over $\mathcal{M}$ or, (ii) by fixing $\mathcal{M}$ to some constant value (and deal with only two free parameters, $\Omega_m$, and $\Omega \Lambda$).

The confidence region ellipses in the $\Omega_m - \Omega \Lambda$ plane are shown in Figure 5 for the full data set. In panel (a) the confidence regions are obtained by marginalizing over $\mathcal{M}$. The best-fit values are found to be $\Omega_m = 0.67 \pm 0.25$, $\Omega \Lambda = 1.24 \pm 0.34$ and $\mathcal{M} = 23.92 \pm 0.05$, as indicated in the panel. For panels (b – d), the confidence contours are obtained by fixing $\mathcal{M}$ to a constant value rather than marginalizing over this parameter. The three frames correspond to the best-fit mean value and two values in the wings of $1 \sigma$ from the mean, respectively. The fixed value of $\mathcal{M}$ is indicated in the panel.

The main conclusions we can draw from this figure are: (i) The results do not change significantly whether we marginalize over $\mathcal{M}$ or whether we use the best-fit value. One can see that the best-fit values of $\Omega_m$ and $\Omega \Lambda$ do not change at all, it is only the spread which decreases slightly when we fix $\mathcal{M}$ to a particular value rather than marginalize it. The two methods give similar results because the probability of $\mathcal{M}$ is sharply peaked (the spread is $\sim 1$ per cent of the mean value). (ii) The results exclude the SCDM model ($\Omega_m = 1$, $\Omega \Lambda = 0$) at a high level of significance, taking into account the maximum uncertainty in $\mathcal{M}$. The importance of this exercise will be evident later when we consider high and low redshift data points separately. (iii) The slanted shape of the probability ellipses show that a particular linear combination of $\Omega_m$ and $\Omega \Lambda$ is selected out by these observations which, in this case, turns out to be $0.82 \Omega_m - 0.57 \Omega \Lambda$. This feature, of course, has nothing to do with supernova data and arises purely because the luminosity distance $Q$ depends strongly on a particular linear combination of $\Omega_m$ and $\Omega \Lambda$. This point is illustrated by plotting the contour of constant luminosity distance, $Q(z = 0.63) = constant$. The coincidence of this line (which roughly corresponds to $Q$ at a redshift in the middle of the data) with the probability ellipses indicates that it is the dependence of the luminosity distance on cosmological parameters which essentially determines the nature of this result. This fact is known in the literature [see e.g., Goobar & Perlmutter (1995)], but we have not seen the actual data presented in this form. We shall use this result in the next section to discuss the possibility of determining the equation of state parameter for the evolving dark energy component. One should note that a complicated likelihood analysis may give confidence contours of a different shape (they are ellipses only because of the assumption of normal distribution of errors here) – however, the shapes of those contours are bound...
Figure 4. Confidence region ellipses in the $\Omega_m - M$ plane for flat models with non-relativistic matter and a cosmological constant. The ellipses corresponding to the 68, 90 and 99 per cent confidence regions are shown. In the left panel, all the 55 data points from Perlmutter et al. (1999) and Reiss et al. (2001) are used. In the middle panel, data points with $z < 0.25$ are used, while in the right panel, we have used data points with $z > 0.25$. We have indicated the best-fit values of $\Omega_m$ and $M$ (with $1\sigma$ errors).

Figure 5. Confidence region ellipses in the $\Omega_m - \Omega_\Lambda$ plane for models with non-relativistic matter and a cosmological constant. The ellipses corresponding to the 68, 90 and 99 per cent confidence regions are shown. (a) The confidence regions are obtained after marginalizing over $M$. The best-fit values of $\Omega_m$ and $M$ are indicated (with $1\sigma$ errors). (b – d) The confidence regions are obtained after fixing $M$ to its best-fit mean value and to the values in the wings of $1\sigma$ limit respectively. The fixed value of $M$ is indicated in the panel. The dashed line corresponds to the flat model ($\Omega_m + \Omega_\Lambda = 1$). We have indicated the value of $\chi^2$ for the best-fit parameters in all the panels. The unbroken slanted line corresponds to the contour of constant luminosity distance, $Q(z = 0.63) = \text{constant}$.
show the degeneracy in \( \Omega_m \) and \( \Omega_\Lambda \) (arising purely from theory) mentioned above. We should also mention that the shapes of the contours of constant \( Q \) at high redshifts are quite different from those at low redshifts – hence more high redshift data can be useful in breaking the degeneracy between \( \Omega_m \) and \( \Omega_\Lambda \).

As discussed in the third point above, the luminosity distance is sensitive only to a particular linear combination of \( \Omega_m \) and \( \Omega_\Lambda \), which is illustrated in Figure 6. In this figure, \( \Omega_m \) and \( \Omega_\Lambda \) are treated as free parameters (within 2\( \sigma \) bounds from the best-fit values mentioned above), but the combination \( 0.82\Omega_m - 0.57\Omega_\Lambda \) is held fixed. It turns out that \( Q(z) \) is not very sensitive to individual values of \( \Omega_m \) and \( \Omega_\Lambda \) at low redshifts when \( 0.82\Omega_m - 0.57\Omega_\Lambda \) is in the range \(-0.14 \pm 0.11 \). This is clear from Figure 6 in which a wide variety of cosmological models are plotted along with the data for a constant value of the above combination. Though some of the models have unacceptable values of \( \Omega_m \) and \( \Omega_\Lambda \) (ruled out by observations of CMBR and structure formation), the supernova measurements alone cannot rule them out. Essentially, the data at \( z < 1 \) is degenerate on the linear combination \( 0.82\Omega_m - 0.57\Omega_\Lambda \). Our analysis of the supernova data shows that the best-fit value for this combination is \( 0.82\Omega_m - 0.57\Omega_\Lambda = -0.14 \pm 0.11 \).

Finally, we comment on the interplay between high and low redshift data for non-flat models. Just as in the case of the flat models, we divide the full data set into low (\( z < 0.25 \)) and high (\( z > 0.25 \)) redshift subsets, and repeated the best-fit analysis. The resulting confidence contours are shown in Figure 7. In panel (a), the confidence contours are plotted for the high redshift data set after marginalizing over \( \Omega_m \). As we stressed before, the SCDM model cannot be ruled out using high redshift data alone. In panels (b–d), we show the corresponding results in which the values of \( \Omega_m \) are fixed (rather than marginalizing over \( \Omega_m \)). Comparing the frames (a–d) of Figure 7 with the frames (a–d) of Figure 5 where the full data set was used, we can draw the following conclusions: (i) The best-fit value for \( \Omega_m \) is now 24.05 \pm 0.38; the 1\( \sigma \) error has now gone up by nearly an order of magnitude compared to the case where the full data was used. Because of this spread, the confidence contours are sensitive to the value of \( \Omega_m \) one uses, unlike the situation where all the data points were used. (ii) Our conclusions will now depend on the exact value of \( \Omega_m \) used. For the best-fit mean value and the lower end of \( \Omega_m \), the high redshift data can rule out the SCDM model [see frames (b) and (d)]. But for the higher end of the allowed 1\( \sigma \) range of \( \Omega_m \), we cannot exclude the SCDM model [see frame (c)]. This essentially shows that it is difficult to rule out models when there are large uncertainties in \( \Omega_m \). Finally, we plot the contours for the low redshift data set after marginalizing over \( \Omega_m \) in panel (e). It is clear that the low redshift data cannot be used to discriminate between cosmological models effectively. This is because \( Q(z) \) at low redshifts is only very weakly dependent on the cosmological parameters. So, even though the acceleration of the universe is a low redshift phenomenon, we cannot reliably determine it using low redshift data alone.

### 4 CONSTRAINTS ON EVOLVING DARK ENERGY

As we have seen in the previous two sections, supernova observations suggest the existence of a non-zero \( \Omega_\Lambda \). However, this very existence of \( \Omega_\Lambda \) raises serious theoretical problems [see Padmanabhan (2002a) for a detailed discussion]. These difficulties have led people to consider the possibility that the dark energy is not just a constant – but is evolving with time. The evolution of the dark energy component can be parametrized by its equation of state parameter \( w_X(z) \), the evolution of the density \( \rho_X \) being given by equation (1). In this section, we shall examine the possibility of constraining \( w_X(z) \) by comparing theoretical models with supernova observations.

One simple, phenomenological, procedure for comparing observations with theory is to parametrize the function \( w_X(z) \) in some suitable form and determine a finite set of parameters in this function using the observations. Theoretical models can then be reduced to a finite set of parameters which can be determined from observations. To illustrate this approach, let us assume that, in the flat universe, \( w_X(z) \) is completely determined by its current value \( (w_0) \) and its rate of change with respect to the scale factor at the present epoch \( (-w_1) \) and is given by the simple form as in equation (14).

Since the cosmological constant corresponds to the case \( w_0 = -1, w_1 = 0 \), it is interesting to see the constraints on \( w_0 \) from supernova data even if we assume \( w_1 = 0 \). In such a study, it should be noted that (i) the acceleration of the universe requires \( w_0 < -1/3 \), and (ii) all values of \( w_0 \) other that \( w_0 = -1 \) leads to a dark energy density which evolves as \((1 + z)^{3(1+w_0)} \). A simple best-fit analysis shows that for a flat model with \( \Omega_m = 0.31 \) and \( \Omega_\Lambda = 0.69 \) the best-fit parameters for flat models, obtained in the previous section, the best-fit value of \( w_0 \) is \(-1.01 \pm 0.12 \) (which is nothing but the conventional cosmological constant). The data clearly rules out models with \( w_0 > -1/3 \) at a high significance level, thereby supporting the existence of a dark energy component with negative pressure.

One can extend the analysis to find the constraints in the \( w_0 - w_1 \) plane. [There is extensive literature on determining dark energy parameters from the current and proposed future supernova observations, e.g., see Maor et al. (2002); Weller & Albrecht (2002); Gerke & Efstathiou (2002) and references therein.] As before, we assume a flat universe with a fixed value of \( \Omega_m \) in the
Figure 7. Confidence region ellipses in the $\Omega_m - \Omega_\Lambda$ plane for models with non-relativistic matter and a cosmological constant. The ellipses corresponding to the 68, 90 and 99 per cent confidence regions are shown. (a) The confidence regions are obtained for high redshift points ($z > 0.25$), after marginalizing over $\cal M$. The best-fit value (with $1\sigma$ error) of $\cal M$ is indicated. (b–d) The confidence regions are obtained for high redshift points ($z > 0.25$), after fixing $\cal M$ to its best-fit value, the upper limit and the lower limit respectively. The fixed value of $\cal M$ is indicated in the panel. (e) The confidence regions are obtained for low redshift points ($z < 0.25$), after marginalizing over $\cal M$. The best-fit value (with $1\sigma$ error) of $\cal M$ is indicated.

Figure 8. Confidence region ellipses in the $w_0 - w_1$ plane for flat models with $\Omega_m = 0.21, 0.31$ and 0.41 respectively, as indicated in the frames. The value of $\cal M$ is chosen to be the best-fit value, which is also indicated. The confidence regions corresponding to the 68, 90 and 99 per cent confidence regions are shown. The square point denotes the equation of state for a universe with a non-evolving dark energy component (the cosmological constant). The unbroken slanted line corresponds to the contour of constant luminosity distance, $Q(z = 0.6) = \text{constant.}$
Figure 9. Sensitivity of $Q(z)$ to the parameters $w_0$ and $w_1$. The curves correspond to the constant values for the fractional change in $Q$ for unit change in $w_0$ (top frames) and $w_1$ (bottom frames) for two redshifts $z = 1$ (left frames) and $z = 3$ (right frames). The value of $Q$ for a particular curve is indicated. A flat cosmological model with $\Omega_m = 0.31$ has been used. The shaded bands across the frames correspond to the regions in which $-1 \leq w_X(z) \leq 0$.

As we have stressed in the previous section and also above, the supernova observations essentially measure $Q(z)$, and hence the accuracy in the determination of $w_0$ and $w_1$ from (both present and planned future) supernova observations will crucially depend on how sensitive $Q$ is to the changes in $w_0$ and $w_1$ [which is clear from the coincidence of the $Q = \text{constant}$ line with the probability ellipses in Figure 8]. A good measure of the sensitivity is provided by the two parameters

$$A(z; w_0, w_1) \equiv \frac{d}{dw_0} \ln [Q(z; w_0, w_1)],$$

$$B(z; w_0, w_1) \equiv \frac{d}{dw_1} \ln [Q(z; w_0, w_1)].$$

Like $Q(z; w_0, w_1)$, the parameters $A$ and $B$ can be computed for a particular cosmological model in a straightforward manner. At any given redshift $z$, we can plot contours of constant $A$ and $B$ in the $w_0 - w_1$ plane. Figure 9 shows the result of such analysis for flat models with $\Omega_m = 0.31$. The two frames on the left are for $z = 1$, while those on the right are for $z = 3$. The top frames show the contours of constant $A$ while the bottom frames show the contours of constant $B$. From the definitions (22) it is clear that $A$ and $B$ can be interpreted as the fractional change in $Q$ for unit change in $w_0$ and $w_1$, respectively. For example, along the line marked $A = 0.2$ (in the top left frame), $Q$ will change by 20 per cent for unit change in $w_0$. It is clear from the two top frames that for most of the interesting region in the $w_0 - w_1$ plane, changing $w_0$ by unity changes $Q$ by about 10 per cent or more. Comparison of the two cases $z = 1$ and $z = 3$ (the two top frames) shows that the sensitivity is higher at high redshift, as to be expected. The shaded bands across the frames correspond to the regions in which $-1 \leq w_X(z) \leq 0$, which is of primary interest in constraining dark energy with negative pressure. One concludes from the above analysis that...
Figure 10. Contours of constant $Q$ in the $\Omega_m - w_1$ (bottom frames) and $\Omega_m - w_0$ (top frames) planes at two redshifts $z = 1$ (left frames) and $z = 3$ (right frames). The value of $w_0$ is fixed to $-1$ in the top frames, while $w_1$ is fixed to 0 in the bottom frames.

Discussion that determining $w_0$ from $Q$ fairly accurately will not be too daunting a task.

The situation, however, is quite different for $w_1$ as illustrated in the two bottom frames. For the same region of the $w_0 - w_1$ plane, $Q$ changes only by a few per cent when $w_1$ changes by unity. This means that $Q$ is much less sensitive to $w_1$ than to $w_0$. It is going to be significantly more difficult to determine a value for $w_1$ from observations of $Q$ (like supernova observations) in the near future.

Comparison of the two cases $z = 1$ and $z = 3$ again shows that the sensitivity is somewhat better at high redshifts, but only marginally.

In the above analysis, we have treated $\Omega_m$ to be a constant. The situation is made worse by the fact that $Q$ also depends on $\Omega_m$. If the variation of $\Omega_m$ mimics that of $w_0$ or $w_1$, then one also needs to determine the sensitivity of $Q$ to $\Omega_m$. Figure 10 shows contours of constant $Q$ in the $\Omega_m - w_1$ (top frames) and $\Omega_m - w_0$ (bottom frames) planes at two redshifts $z = 1$ (left frames) and $z = 3$ (right frames). The two top frames show that if one varies the value of $\Omega_m$ in the allowed range, say, $(0.2 - 0.4)$, one can move along the curve of constant $Q$ and induce a fairly large variation ($\sim 1 - 2$) in $w_1$. In other words, large changes in $w_1$ can be easily compensated by small changes in $\Omega_m$ while maintaining the same value for $Q$ at a given redshift. This shows that the uncertainty in $\Omega_m$ introduces further difficulties in determining $w_1$ accurately from measurements of $Q$. The two bottom frames show that the situation is better for $w_0$. The curves are much less steep and hence varying $\Omega_m$ does not induce large variations in $w_0$. We once again are led to the conclusion that unambiguous determination of $w_1$ from data will be quite difficult. It appears that supernova observations may not be of great help in ruling out (non-evolving) cosmological constant as the major dark energy component.

5 DISCUSSION

The supernova data on the whole, rules out non-accelerating models with very high significance level, as noted and stressed by various authors (Riess et al. 1998; Perlmutter et al. 1999; Riess 2000). It is thus more or less conclusive that we have to look for some form of matter which has negative pressure. However, it is interesting to note that if we divide the data set into high and low redshift subsets, neither of the subsets are able to rule out the decelerating models – it is only the interplay between the high and low redshift data which implies the result quoted above. This analysis indirectly stresses the importance of any evolutionary effects. Any possible evolution in the absolute magnitude of the supernovae, if detected, might allow the decelerating models to be consistent with the data. However, our understanding of the complicated physical processes in the supernovae is definitely not well enough to discuss the possibility of detecting supernova evolution in near future. Although the intrinsic evolution of the SN is an interesting problem, addressing these issues is beyond the scope of this paper.

The supernova data available today is not enough to fix the parameters $\Omega_m$ and $\Omega_\Lambda$ independently. Instead, a particular combination of the two parameters is chosen out by the data. This degeneracy has to do with the dependence of the luminosity distance on these parameters. Since supernova observations essentially measure the luminosity distance $Q(z)$, the accuracy in the determina-
tion of various parameters from supernova observations will crucially depend on how sensitive $Q$ is to the changes in those parameters. This analysis can be done entirely from theory [using Fisher information matrix; see for example Efstathiou (1999)] and, in principle, can provide nice handle on which combinations of parameters are expected to be constrained from the supernova data.

The key issue regarding dark energy is whether it is a constant or whether it is evolving with time. In particular, one is interested to determine the evolution of the equation of state, $w_X$ of the dark energy component. To address this question, we have taken a simple phenomenological model for $w_X$ having the form $w_X(a) = w_0 - w_1 (a - 1)$. The sensitivity of the luminosity distance (and hence, the supernova data) on $w_0$ and $w_1$ can be measured by the two parameters ($A$ and $B$) introduced in the paper, which are essentially the fractional changes in $Q(z)$ for unit change in $w_0$ and $w_1$, respectively. The parameters $A$ and $B$ can be obtained entirely from theory and the conclusions drawn from them are valid for current as well as future supernova observations. We find that $Q(z)$ is quite sensitive to $w_0$ – hence, one can constrain the current value of $w_X$ quite well. However, $Q(z)$ is comparatively insensitive to $w_1$, thus determining the evolution of $w_X$ will be a difficult task. The situation is further worsened when we take the uncertainties in $\Omega_m$ into account. Our results are consistent with other statistical analyses done by Astier (2000), Maor, Brustein, & Steinhardt (2001), Weller & Albrecht (2001), Maor et al. (2002), Weller & Albrecht (2002). They have shown that it is possible to constrain the present value of $\omega_m$ with future missions like “SuperNovae Acceleration Probe” (SNAP); however, constraining the evolution of $w_X$ will not be easy using SNAP, particularly when $\Omega_m$ is not known to a high accuracy.

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APPENDIX A: DATA ANALYSIS

In this section, we shall outline the method used for obtaining the best-fit (theoretical) parameters from observational data, and draw the corresponding confidence contours. For simplicity, we shall assume that the measurement errors are normally distributed. Suppose we have $M$ observational data points denoted by $m_{obs}(z_i); i = 1, ..., M$ with corresponding errors $\sigma_m(z_i)$. For a given theoretical model $m_{th}(z; c_\alpha)$ with $\nu$ free parameters $c_\alpha; \alpha = 1, ..., \nu$, one can construct the quantity

$$\chi^2(c_\alpha) = \sum_{i=1}^{M} \left[ \frac{m_{obs}(z_i) - m_{th}(z_i; c_\alpha)}{\sigma_m(z_i)} \right]^2$$

(A1)

The best-fit parameters $\bar{c}_\alpha$ are obtained by minimizing the above quantity

$$\frac{\partial \chi^2}{\partial c_\alpha} \bigg|_{c_\alpha = \bar{c}_\alpha} = 0$$

(A2)

using the Levenberg-Marquardt method (Press et al. 1992). One can show that the quantity $\Delta \chi^2 = \chi^2(\bar{c}_\alpha) - \chi^2(\bar{c})$ follows a chi-square distribution with $\nu$ degrees of freedom.

To obtain the confidence intervals, let us first define the curvature matrix

$$A_{\alpha\beta} = \frac{1}{2} \left[ \frac{\partial^2 \chi^2}{\partial c_\alpha \partial c_\beta} \right]_{c_\alpha = \bar{c}_\alpha}$$

(A3)

The covariance matrix $C_{\alpha\beta}$ is simply the inverse of the curvature matrix. The probability distribution of the parameters is given by

$$P(c_\alpha) = \text{const} \times \exp \left[ -\frac{1}{2} \sum_{\alpha,\beta=1}^{\nu} (c_\alpha - \bar{c}_\alpha) A_{\alpha\beta} (c_\beta - \bar{c}_\beta) \right]$$

(A4)

Suppose we are interested in the confidence regions for a subset of, say, $\nu'$ parameters ($\nu' \leq \nu$). In that case, the the regions are obtained by simply marginalizing the probability distribution over the rest $\nu - \nu'$ parameters [which is equivalent to integrating the probability distribution (A4) over the rest $\nu - \nu'$ parameters]. This can be done simply by taking the full $\nu \times \nu$ covariance matrix $C_{\alpha\beta}$ and copying the intersection of the $\nu'$ rows and columns corresponding to the parameters of interest into a $\nu' \times \nu'$ matrix $C'_{\alpha\beta}$. The inverse of this will give the corresponding curvature matrix $A'_{\alpha\beta}$ (Press et al. 1992). The marginalized probability distribution will be simply

$$P'(c_\alpha) = \text{const} \times \exp \left[ -\frac{1}{2} \sum_{\alpha,\beta=1}^{\nu'} (c_\alpha - \bar{c}_\alpha) A'_{\alpha\beta} (c_\beta - \bar{c}_\beta) \right]$$

(A5)

Next, one needs to find the quantity $\Delta \chi^2_{\nu'}$, such that the probability of a chi-square variable with $\nu'$ degrees of freedom being less
than $\Delta \chi^2_p$ is p, where p is the desired confidence limit (e.g., 0.68 or 0.95). This can be obtained by solving the equation

$$p = \frac{1}{2^{\nu'/2}\Gamma(\nu'/2)} \int_0^{\Delta \chi^2_p} \, du \, u^{(\nu'/2) - 1} e^{-u/2}$$

(A6)

where $\Gamma(\cdot)$ is the gamma function. For a given p and a set of $\nu'$ parameters, the equation for the boundary of the confidence region in the $\nu'$-dimensional space is given by the equation

$$\Delta \chi^2_p = \sum_{\alpha,\beta=1}^{\nu'} (e_{\alpha} - \bar{e}_{\alpha}) A_{\alpha,\beta} (e_{\beta} - \bar{e}_{\beta})$$

(A7)

**APPENDIX B: DETERMINATION OF THE HUBBLE PARAMETER $H(z)$ FROM SUPERNOVA OBSERVATIONS**

The observational data used in this paper can be fitted by the function of simple form

$$m_{\text{fit}}(z) = a_1 + 5 \log_{10} \left( \frac{z(1 + a_2 z)}{1 + a_3 z} \right).$$

(B1)

where the parameters are

$$a_1 = 23.95 \pm 0.05; \ a_2 = 2.00 \pm 1.18; \ a_3 = 1.03 \pm 0.88$$

(B2)

The $\chi^2$ per degree of freedom for the best-fit values of the parameters is found to be 1.05, which shows the fit is reasonably good. The difference between the data points and the fit is shown in Figure B1. We can then represent the luminosity distance obtained from the data by the function

$$Q_{\text{fit}}(z) = 10^{0.2 m_{\text{fit}}(z) - K}$$

(B3)

Note that one needs to fix the value of $\mathcal{M}$ to obtain the function $Q_{\text{fit}}(z)$. For flat models, the value of $\mathcal{M}$ is well-constrained (the error being ~ 1 per cent) – hence we can use the best-fit value of $\mathcal{M} = 23.95$ to obtain

$$Q_{\text{fit}}(z) = \frac{z(1 + 2.00 z)}{1 + 1.03 z}$$

(B4)

For flat models, it straightforward to obtain the Hubble parameter from $Q(z)$. In particular, we are interested in the quantity

$$H_0^{-1} \dot{a}(z) = \left[ (1 + z) \frac{d}{dz} \left( \frac{Q(z)}{1 + z} \right) \right]^{-1}$$

(B5)

which will enable us to plot the data points in the $\dot{a} - a$ plane. The determination of the corresponding error-bars is a non-trivial exercise. In this paper, the errors are obtained from the relation

$$H_0^{-1} \dot{a}(z)(1 + \epsilon_a(z)) = \left[ (1 + z) \frac{d}{dz} \left( \frac{Q(z)(1 + \epsilon_Q(z))}{1 + z} \right) \right]^{-1}$$

(B6)

where $\epsilon_a$ and $\epsilon_Q$ denote the fractional error in $\dot{a}$ and $Q$, respectively. From equation (18), we have $\epsilon_Q(z) = 0.2 \ln 10 \sigma_m(z)$, where $\sigma_m(z)$ is the total uncertainty in the observed magnitude. Since there is no systematic evolution of the observed $\sigma_m(z)$ (see Figure B1), we have, to the lowest approximation, the fractional error in $\dot{a}$

$$\epsilon_a(z) = \epsilon_Q(z) = 0.2 \ln 10 \sigma_m(z).$$

(B7)

Thus, one can plot the quantity $\dot{a}(z)$ for the observational data with error-bars given by the above equation.

Note that the analysis assumes that the observed errors in the measurement of $z$ (or, equivalently $a$) are negligible. Strictly speaking, the errors in $\dot{a}$ calculated above are strict lower limits – they can be slightly higher because of the errors arising from (i) the fitting function and (ii) any systematic evolution of the observed $\sigma_m(z)$. However, these effects are unlikely to affect any of the conclusions we have drawn in this paper.

Figure B1. The difference between the actual data points and the fit (B1) to the data.