Communication Requirements for Generating Correlated Random Variables

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Abstract—Two familiar notions of correlation are rediscovered as extreme operating points for simulating a discrete memoryless channel, in which a channel output is generated based only on a description of the channel input. Wyner’s “common information” coincides with the minimum description rate needed. However, when common randomness independent of the input is available, the necessary description rate reduces to Shannon’s mutual information. This work characterizes the optimal tradeoff between the amount of common randomness used and the required rate of description.

I. INTRODUCTION

What is the intrinsic connection between correlated random variables? How much interaction is necessary to create correlation?

Many fruitful efforts have been made to quantify correlation between two random variables. Each quantity is justified by the operational questions that it answers. Covariance dictates the mean squared error in linear estimation. Shannon’s mutual information is the descriptive savings from side information in lossless source coding and the additional growth rate of wealth due to side information in investing. Gács and Körner’s common information [1] is the number of common random bits that can be extracted from correlated random variables. It is less than mutual information. Wyner’s common information [2] is the number of common random bits needed to generate correlated random variables and is greater than mutual information.

This work provides a fresh look at two of these quantities — mutual information and Wyner’s common information (herein simply “common information”). Both are extreme points of the channel simulation problem, introduced as follows: An observer (encoder) of an i.i.d. source $X_1, X_2, ...$ describes the sequence to a distant random number generator (decoder) that produces $Y_1, Y_2, ...$ (see Figure 1). What is the minimum rate of description needed to achieve a joint distribution that is statistically indistinguishable (as measured by total variation) from the distribution induced by putting the source through a memoryless channel?

Channel simulation is a form of random number generation. The variables $X^n$ come from an external source and $Y^n$ are generated to be correlated with $X^n$. The channel simulation is successful if the total variation between the resulting distribution of $(X^n, Y^n)$ and the i.i.d. distribution that would result from passing $X^n$ through a memoryless channel is small. This is a strong requirement. It’s stricter than the requirement that $(X^n, Y^n)$ be jointly typical as in the coordinated action work of Cover and Permuter [3]. This total variation requirement means that any hypothesis test that a statistician comes up with to determine whether $X^n$ was passed through a real memoryless channel or the channel simulator will be virtually useless.

Wyner’s result implies that in order to generate $X^n$ and $Y^n$ separately as an i.i.d. source pair they must share bits at a rate of at least the common information $C(X; Y)$ of the joint distribution. In the channel simulation problem these shared bits come in the form of the description of $X^n$. However, the “reverse Shannon theorem” of Bennett and Shor [4] suggests that a description rate of the mutual information $I(X; Y)$ of the joint distribution is all that is needed to successfully simulate a channel. How can we resolve this apparent contradiction?

The work of Bennett and Shor assumes that common random bits, or common randomness, independent of the source $X^n$ are available to the encoder and decoder. In that setting, the common randomness provides a second connection between the source $X^n$ and output $Y^n$.

1To achieve channel simulation with a rate as low as the common information one must change Wyner’s relative entropy requirement in [2] to a total variation requirement as used in this work.
Y^n$, in addition to the description of $X^n$. Remarkably, even though it is independent from the source $X^n$, the common randomness assists in generating correlated random numbers and allows for description rates smaller than the common information $C(X; Y)$.

In this work, we characterize the tradeoff between the rate of available common randomness and the required description rate for simulating a discrete memoryless channel for a fixed input distribution, as in Figure 2. Indeed, the tradeoff region of Section III confirms the two extreme cases. If the encoder and decoder are provided with enough common randomness, sending $I(X; Y)$ bits per symbol suffices. On the other hand, in the absence of common randomness one must spend $C(X; Y)$ bits per symbol.

This result has implications in cooperative game theory, reminiscent of the framework investigated in [5]. Suppose a team shares the same payoff in a repeated game setting. An opponent tries to anticipate and exploit patterns in the team’s combined actions, but a secure line of communication is available to help them coordinate. Of course, each player could communicate his randomized actions to the other players, but this is an excessive use of communication. A memoryless channel is a useful way to coordinate their random actions. Thus, common information is found in Section VII to be the significant quantity in this situation.

II. PRELIMINARIES AND PROBLEM DEFINITION

A. Notation

We represent random variables as capital letters, $X$, and their alphabets are written in script, $\mathcal{X}$. Sequences, $X_1, ..., X_n$, are indicated with a superscript $X^n$. Distribution functions, $p_X(x)$, are usually abbreviated as $p(x)$ when there is no confusion.

Accentted variables, $\hat{X}$, indicate different variables for each accent, but their alphabets are all the same, $\mathcal{X}$. Similarly, distribution functions written with an accent or different letter, such as $p(x)$ versus $\hat{p}(x)$, represent different distributions.

Markov chains, satisfying $p(x, y, z) = p(x, y)p(z|y)$, are represented with dashes, $X - Y - Z$.

(Wyner’s) common information:

$$C(X; Y) \triangleq \min_{X \rightarrow U \rightarrow Y} I(X, Y; U).$$

Conditional common information:

$$C(X; Y|W) \triangleq \min_{X \rightarrow U \rightarrow Y} I(X, Y; U|W).$$

Total variation distance:

$$||p - q||_1 \triangleq \frac{1}{2} \sum_x |p(x) - q(x)|.$$

B. Problem Specific Definitions

A source $X^n$ is distributed i.i.d. according to $\bar{p}(x)$. A description of the source at rate $R_1$ is represented by $I \in \{1, ..., 2^{nR_1}\}$. A random variable $J$, uniformly distributed on $\{1, ..., 2^{nR_2}\}$ and independent of $X^n$, represents the common random bits at rate $R_2$ known at both the encoder and decoder. The decoder generates a channel output $Y^n$ based only on $I$ and $J$.

The channel being simulated has a the conditional distribution $q(y|x)$, thus the desired joint distribution is $\bar{p}(x)q(y|x)$.

Definition 1: A $(2^{nR_1}, 2^{nR_2}, n)$ channel simulation code consists of a randomized encoding function,

$$F_n : \mathcal{X}^n \times \{1, 2, ..., 2^{nR_1}\} \rightarrow \{1, 2, ..., 2^{nR_1}\},$$

and a randomized decoding function,

$$G_n : \{1, 2, ..., 2^{nR_1}\} \times \{1, 2, ..., 2^{nR_2}\} \rightarrow \mathcal{Y}^n.$$ 

The description $I$ equals $F_n(X^n, J)$, and the channel output $Y^n$ equals $G_n(I, J)$.

Since randomized functions are specified by conditional probability distributions, it is equivalent to say that a $(2^{nR_1}, 2^{nR_2}, n)$ channel simulation code consists of a conditional probability mass function $p(i, y^n|x^n, j)$ with the properties that $p(y^n|i, j, x^n) = p(y^n|j, i)$, $|I| = 2^{nR_1}$, and $|J| = 2^{nR_2}$.

The induced joint distribution of a $(2^{nR_1}, 2^{nR_2}, n)$ channel simulation code is the joint distribution on the quadruple $(X^n, Y^n, I, J)$. In other words, it is the probability mass function,

$$p(x^n, y^n, i, j) = p(i, y^n|x^n, j)p(x^n, j), \quad (1)$$

where $p(x^n, j) = p(j) \prod_{k=1}^n \bar{p}(x_k)$ by construction.

Definition 2: A sequence of $(2^{nR_1}, 2^{nR_2}, n)$ channel simulation codes for $n = 1, 2, ...$ is said to achieve $q(y|x)$ if the induced joint distributions have marginal distributions $p(x^n, y^n)$ that satisfy

$$\lim_{n \to \infty} \left\| p(x^n, y^n) - \prod_{k=1}^n \bar{p}(x_k)q(y_k|x_k) \right\|_1 = 0.$$

Definition 3: A rate pair $(R_1, R_2)$ is said to be achievable if there exists a sequence of $(2^{nR_1}, 2^{nR_2}, n)$ channel simulation codes that achieves $q(y|x)$.

Definition 4: The simulation rate region is the closure of achievable rate pairs $(R_1, R_2)$. 



\[\text{Fig. 2. A discrete-memoryless channel is simulated by two separate processors, } F \text{ and } G. \text{ The first processor, } F \text{, observes } X \text{ and common randomness independent of } X \text{ at rate } R_2. \text{ The second processor, } G \text{, generates } Y \text{ based on the common randomness and a message at rate } R_1 \text{ from } F.\]
The mutual information terms in (2) become
\[ D \triangleq \{ (R_1, R_2) \in \mathbb{R}^2 : \exists p(x, y, u) \in D \text{ s.t.} \]
\[ R_1 \geq I(X; U), \]
\[ R_1 + R_2 \geq I(X, Y; U) \}, \tag{2} \]
where
\[ D \triangleq \{ p(x, y, u) : (X, Y) \sim \hat{p}(x)q(y|x), \]
\[ X - U - Y \text{ form a Markov chain,} \]
\[ |U| \leq |X||Y| + 1 \}. \tag{3} \]

IV. OBSERVATIONS AND EXAMPLES

Two extreme points of the simulation rate region \( S \) fall directly from its definition. If \( R_2 = 0 \), the second inequality in (2) dominates. Thus, the minimum rate \( R_1 \) is the common information \( C(X; Y) \). This coincides with the intuition provided by Wyner’s result in [2]. At the other extreme, using the data processing inequality on the first inequality of (2) yields \( R_1 \geq I(X; Y) \) no matter how much common randomness is available, and this is achieved when \( R_2 \geq H(Y|X) \). Source coding results and the coordinated action work of Cover and Permuter in [3] illustrate that with a description rate of \( I(X; Y) \) we can create a codebook of output sequences in such a way that we’ll likely be able to find a jointly typical output sequence for each input sequence from the source. Consequently, we can then randomize the codebook using common randomness to actually simulate the channel, as Bennett and Shor proved in [4].

A. Binary Erasure Channel

For a Bernoulli-half source \( X \), let us demonstrate the simulation rate region for the binary erasure channel. \( Y \) is an erasure with probability \( P_e \) and is equal to \( X \) otherwise. The distributions in \( D \) that produce the boundary of the simulation rate region are formed by cascading two binary erasure channels as shown in Figure 3 where
\[ p_2 \in \left[ 0, \min \left\{ \frac{1}{2}, P_e \right\} \right], \]
\[ p_1 = 1 - \frac{1 - P_e}{1 - p_2}. \]

The mutual information terms in (2) become
\[ I(X; U) = 1 - p_1, \]
\[ I(X, Y; U) = h(P_e) + (1 - p_1)(1 - h(p_2)), \]
where \( h \) is the binary entropy function.

\( R_2 \) doesn’t necessarily have to be as large as \( H(Y|X) \) for \( (I(X; Y), R_2) \) to be in the simulation rate region.

Figure 4 shows the boundary of the simulation rate region for an erasure channel with erasure probability \( P_e = 0.75 \) and a Bernoulli-half input, where \( R_1 \) is the description rate and \( R_2 \) is the rate of common randomness. Without common randomness, a description rate of \( C(X; Y) \) is required to simulate the channel. With unlimited common randomness, a description rate of \( I(X; Y) \) suffices.

V. SKETCH OF CONVERSE

Let \((R_1, R_2)\) be an achievable rate pair. Then for each \( \epsilon \in (0, 1/4) \) there exists a \((2^{nR_1}, 2^{nR_2}, n)\) channel simulation code with an induced joint distribution \( p(x^n, y^n, i, j) \) such that
\[ \left\| p(x^n, y^n) - \prod_{k=1}^n \hat{p}(x_k)q(y_k|x_k) \right\|_1 < \epsilon. \]

Let the random variable \( K \) be uniformly distributed over the set \( \{1, \ldots, n\} \). The variable \( K \) will serve as a random time index.

A. Entropy Bounds

The joint distribution of the sequences \((X^n, Y^n)\) is close in total variation to an i.i.d. distribution, so we can
extend Lemma 2.7 of [6] to obtain two bounds:

\[
H(X^n, Y^n) - \sum_{k=1}^{n} H(X_k, Y_k) \leq n g(\epsilon),
\]

\[
I(X^n, Y^n; K) \leq n g(\epsilon),
\]

where

\[
g(\epsilon) \triangleq 4e \left( \log |\mathcal{X}| + \log |\mathcal{Y}| + \log \frac{1}{\epsilon} \right).
\]

Notice that \( \lim_{\epsilon \to 0} g(\epsilon) = 0. \)

**B. Epsilon Rate Region**

Define an epsilon rate region,

\[
S_\epsilon \triangleq \{(R_1, R_2) \in \mathbb{R}^2 : \exists p(x, y, u) \in D_\epsilon \text{ s.t.} \quad R_1 \geq I(X; U) - 2g(\epsilon), \quad R_1 + R_2 \geq I(X, Y; U) - 2g(\epsilon)\},
\]

where

\[
D_\epsilon \triangleq \{p(x, y, u) : ||p(x, y) - \bar{p}(x)q(y|x)||_1 < \epsilon, \quad X - U - Y \text{ form a Markov chain,} \quad |U| \leq |X||Y| + 1\}.
\]

**Lemma 5.1:**

\((R_1, R_2) \in S_\epsilon.\)

**Proof:** We use familiar information theoretic inequalities, and the fact that \(X^n\) and \(J\) are independent, to bound \(R_1\) and the sum rate \(R_1 + R_2.\)

\[
nR_1 \geq H(I) \geq H(I|J) \geq I(X^n; I|J) = I(X^n; I, J).
\]

\[
n(R_1 + R_2) \geq H(I, J) \geq I(X^n, Y^n; I, J).
\]

We then lower bound the r.h.s. of (8) and (9) using similar steps. Here we proceed from (9).

\[
I(X^n; Y^n; I, J) = H(X^n, Y^n) - H(X^n, Y^n|I, J)
\]

\[
\geq H(X^n, Y^n) - \sum_{k=1}^{n} H(X_k, Y_k|I, J)
\]

\[
\geq \sum_{k=1}^{n} I(X_k, Y_k; I, J) - ng(\epsilon)
\]

\[
= nI(X_k, Y_k; I, J|K) - ng(\epsilon)
\]

\[
\geq nI(X_k, Y_k; I, J, K) - 2ng(\epsilon).
\]

The second inequality comes from (4), and the last inequality comes from (5).

The joint distribution of the pair \((X_K, Y_K)\) can be shown to satisfy the total variation constraint in (7). Finally, we acknowledge the Markovity of the triple \(X_K - (I, J, K) - Y_K\) to complete the proof of the lemma. (The cardinality bound of \(U\) in (7) is shown to be satisfiable via a generalized Caratheodory theorem.)

**C. Lower semi-continuity**

The epsilon rate regions decrease to the simulation rate region as epsilon decreases to zero.

**Lemma 5.2:**

\[
\bigcap_{\epsilon \in (0, 1/2)} S_\epsilon \subset S.
\]

**VI. SKETCH OF ACHIEVABILITY**

**A. Resolvability**

One key tool for the achievability proof is summarized in Lemma 6.1. This lemma is implied by the resolvability work of Han and Verdù in [7], but the concept was first introduced by Wyner in Theorem 6.3 of [2].

**Lemma 6.1:** For any discrete distribution \(p(u, v)\) and each \(n\), let \(C^{(n)} = \{U^n(m)\}_{m=1}^{2n}\) be a “codebook” of sequences each independently drawn according to \(\prod_{k=1}^{n} p_U(u_k).\)

Then for \(R > I(V; U),\)

\[
\lim_{n \to \infty} \mathbb{E} \left| Q(v^n) - \sum_{k=1}^{n} p_V(v_k) \right|_1 = 0,
\]

where the expectation is with respect to the randomly constructed codebooks \(C^{(n)}.\)

**B. Existence of Achievable Codes**

Assume that \((R_1, R_2)\) is in the interior of \(S\). Then there exists a distribution \(p^*(x, y, u) \in D\) such that \(R_1 > I(X; U)\) and \(R_1 + R_2 > I(X, Y; U).\)

For each \(n\), let \((I, J)\) be uniformly distributed on \(\{1, \ldots, 2nR_1\} \times \{1, \ldots, 2nR_2\}.\) We apply Lemma 6.1 twice, once with \(V = (X, Y)\) and again with \(V = X,\) to assert that there exists a sequence of “codebooks” \(C^{(n)} = \{U^n(i, j)\}_{i, j \in \mathcal{I} \times \mathcal{J}}, n = 1, 2, \ldots\) with the properties

\[
\lim_{n \to \infty} \left\| Q(x^n, y^n) - \prod_{k=1}^{n} p_{X,Y}^*(x_k, y_k) \right\|_1 = 0,
\]

\[
\lim_{n \to \infty} \left\| Q(x^n, j) - p(j) \prod_{k=1}^{n} p_X^*(x_k) \right\|_1 = 0,
\]

where \(Q(x^n, y^n)\) and \(Q(x^n, j)\) are marginal distributions derived from the joint distribution

\[
Q(x^n, y^n, i, j) = p(i, j) \prod_{k=1}^{n} p_{X,Y}^*(x_k, y_k U_k(i, j)).
\]
In an indirect way, we’ve constructed a sequence of joint distributions \( Q(x^n, y^n, i, j) \) from which we can derive channel simulation codes that achieve \( q(y|x) \). The Markovity of \( p^* \) implies the Markov property \( Q(x^n, y^n|i, j) = Q(x^n|i, j)Q(y^n|i, j) \). Let
\[
\hat{p}(i|x^n,j) = Q(i|x^n,j),
\]
\[
\hat{p}(y^n|i, j) = Q(y^n|i, j).
\]
Considering (10) and (11) with the properties of total variation and \( p^* \) in mind, it can be shown that \( \hat{p}(i, y^n|x^n, j) = \hat{p}(i|x^n, j)\hat{p}(y^n|i, j) \) is a sequence of channel simulation codes that achieves \( q(y|x) \).

C. Comment on Achievability Scheme

This channel simulation scheme requires randomization at both the encoder and decoder. In essence, a codebook of independently drawn \( U^n \) sequences is overpopulated so that the encoder can choose one randomly from many that are jointly typical with \( X^n \). The decoder then randomly generates \( Y^n \) conditioned on \( U^n \).

VII. GAME THEORY

Our framework finds motivation in a game theoretic setting. Consider a zero-sum repeated game between two teams. Team A consists of two players who on the \( i \)th iteration take actions \( X_i \in \mathcal{X} \) and \( Y_i \in \mathcal{Y} \). The opponents on team B take combined action \( Z_i \in \mathcal{Z} \). All action spaces \( \mathcal{X}, \mathcal{Y}, \mathcal{Z} \) are finite. The payoff for team A at each iteration is a time-invariant finite function \( \Pi(X_i, Y_i, Z_i) \) and is the loss for team B. Each team wishes to maximize its time-averaged expected payoff.

Assume that team A plays conservatively, attempting to maximize the expected payoff for the worst-case actions of team B. Then the payoff at the \( i \)th iteration is
\[
\Theta_i \triangleq \min_{z \in \mathcal{Z}} \mathbb{E}[\Pi(X_i, Y_i, z)|X^{i-1}, Y^{i-1}] .
\]
Clearly, (12) could be maximized by finding an optimal mixed strategy \( p^*(x, y) \) that maximizes \( \min_{z \in \mathcal{Z}} \mathbb{E}[\Pi(X, Y, z)] \) and choosing independent actions each iteration. This would correspond to the minimax strategy. However, now we introduce a new constraint: The players on team A have a limited secure channel of communication. Player 1, who chooses the actions \( X^n \), communicates at rate \( R \) to Player 2, who chooses \( Y^n \).

Let \( U \) be the message passed from Player 1 to Player 2. We say a rate \( R \) is achievable for payoff \( \Theta \) if there exists a sequence of random variable triples \( (X^n, Y^n, U) \) that each form Markov chains \( X^n \rightarrow U \rightarrow Y^n \) and such that \( |U| \leq 2^{nR} \) and
\[
\lim_{n \rightarrow \infty} \mathbb{E} \left[ \frac{1}{n} \sum_{i=1}^{n} \Theta_i \right] > \Theta.
\]

Let \( R(\Theta) \) be the infimum of achievable rates for payoff \( \Theta \). We claim that \( R(\Theta) \) is the least average common information of all combinations of strategies that achieve average payoff \( \Theta \). Define,
\[
R_0(\Theta) \triangleq \min C(X; Y|W)
\]
\[
s.t. \mathbb{E} \left[ \min_{z \in \mathcal{Z}} \mathbb{E}[\Pi(X, Y, z)|W] \right] \geq \Theta.
\]

Theorem 7.1:
\[
R(\Theta) = R_0(\Theta).
\]

Converse Sketch:
The important elements of the converse are the inequalities
\[
n(R(\Theta) + \epsilon) > H(U)
\]
\[
\geq I(X^n, Y^n; U)
\]
\[
= \sum_{i=1}^{n} I(X_i, Y_i; U|X^{i-1}, Y^{i-1})
\]
\[
= nI(X_{K}, Y_{K}; U|X^{K-1}, Y^{K-1}, K),
\]
for all \( \epsilon > 0 \), where \( K \) is uniformly distributed on \( \{1, ..., n\} \). Now identify the tuple \( (X^{K-1}, Y^{K-1}, K) \) as the auxiliary random variable \( W \).

Achievability Comment:
The random variable \( W \) serves as a time sharing variable to combine strategies of high and low correlation.

VIII. ACKNOWLEDGMENT

The author would like to thank his advisor, Tom Cover, for encouraging the study of coordination via communication, and Young-Han Kim for his suggestions and encouragement. This work is supported by the National Science Foundation through grants CCF-0515303 and CCF-0635318.

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