A probabilistic framework for multi-view feature learning with many-to-many associations via neural networks

Akifumi Okuno 1 2  Tetsuya Hada 3  Hidetoshi Shimodaira 1 2

Abstract

A simple framework Probabilistic Multi-view Graph Embedding (PMvGE) is proposed for multi-view feature learning with many-to-many associations so that it generalizes various existing multi-view methods. PMvGE is a probabilistic model for predicting new associations via graph embedding of the nodes of data vectors with links of their associations. Multi-view data vectors with many-to-many associations are transformed by neural networks to feature vectors in a shared space, and the probability of new association between two data vectors is modeled by the inner product of their feature vectors. While existing multi-view feature learning techniques can treat only either of many-to-many association or non-linear transformation, PMvGE can treat both simultaneously. By combining Mercer’s theorem and the universal approximation theorem, we prove that PMvGE learns a wide class of similarity measures across views. Our likelihood-based estimator enables efficient computation of non-linear transformations of data vectors in large-scale datasets by minibatch SGD, and numerical experiments illustrate that PMvGE outperforms existing multi-view methods.

Table 1. Comparison of PMvGE with existing multi-view / graph-embedding methods. (Nv): Number of views, (MM): Many-to-many, (NL): Non-linear, (Ind): Inductive, (Lik): Likelihood-based. PMvGE has all the properties. Nv = D represents that the method can deal with arbitrary number of views.

| Method   | (Nv) | (MM) | (NL) | (Ind) | (Lik) |
|----------|------|------|------|-------|-------|
| CCA      | 2    | ✓    | ✓    |       | ✓     |
| DCCA     | 2    | ✓    | ✓    |       | ✓     |
| MCCA     | D    | ✓    |       |       | ✓     |
| SGE      | 0    | ✓    |       |       | ✓     |
| LINE     | 0    | ✓    | ✓    |       | ✓     |
| LPP      | 1    | ✓    | ✓    |       | ✓     |
| CvGE     | 2    | ✓    | ✓    |       | ✓     |
| CDMCA    | D    | ✓    | ✓    |       | ✓     |
| DeepWalk | 0    | ✓    | ✓    |       | ✓     |
| SBM      | 1    | ✓    | ✓    |       | ✓     |
| GCN      | 1    | ✓    | ✓    |       | ✓     |
| GraphSAGE| 1    | ✓    | ✓    | ✓     | ✓     |
| IDW      | 1    | ✓    | ✓    | ✓     | ✓     |
| PMvGE    | D    | ✓    | ✓    | ✓     | ✓     |

“views”, and their dimensions may be different depending on the view. Typical examples are data vectors of images, text tags, and user attributes available in Social Networking Services (SNS). However, we cannot apply standard data analysis methods, such as clustering, to multi-view data vectors, because data vectors from different views, say, images and text tags, are not directly compared with each other. In this paper, we work on multi-view Feature Learning for transforming the data vectors from all the views into new vector representations called “feature vectors” in a shared euclidean subspace.

One of the best known approaches to multi-view feature learning is Canonical Correlation Analysis (Hotelling, 1936, CCA) for two-views, and Multiset CCA (Kettenring, 1971, MCCA) for many views. CCA considers pairs of related data vectors \( \{(x_{i}^{(1)}, x_{i}^{(2)})\}_{i=1}^{n} \subset \mathbb{R}^{p_1} \times \mathbb{R}^{p_2} \). For instance, \( x_{i}^{(1)} \in \mathbb{R}^{p_1} \) may represent an image, and \( x_{i}^{(2)} \in \mathbb{R}^{p_2} \) may represent a text tag. Their dimension \( p_1 \) and \( p_2 \) may be different. CCA finds linear transformation matrices \( A^{(1)}, A^{(2)} \) so that the sum of inner products \( \sum_{i=1}^{n} (A^{(1)} x_{i}^{(1)} A^{(2)} x_{i}^{(2)})^{T} \) is maximized under a variance constraint. The obtained linear transformations compute feature vectors \( y_{i}^{(1)} := A^{(1)} x_{i}^{(1)}, y_{i}^{(2)} := A^{(2)} x_{i}^{(2)} \). The obtained multi-view embeddings enable efficient computation of non-linear transformations of data vectors in large-scale datasets by minibatch SGD.

1Graduate School of Informatics, Kyoto University, Kyoto, Japan 2RIKEN Center for Advanced Intelligence Project (AIP), Tokyo, Japan 3Recruit Technologies Co., Ltd., Tokyo, Japan. Correspondence to: Akifumi Okuno <okuno@sys.i.kyoto-u.ac.jp>.

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\(A^{(2)\top} x^{(2)}_i \in \mathbb{R}^K\) where \(K \leq \min\{p_1, p_2\}\) is the dimension of the shared space of feature vectors from the two views. However, the linear transformations may not capture the underlying structure of real-world datasets due to its simplicity.

To enhance the expressiveness of transformations, CCA has been further extended to non-linear settings, as Kernel CCA (Lai and Fyfe, 2000) and Deep CCA (Andrew et al., 2013; Wang et al., 2016, DCCA) which incorporate kernel methods and neural networks to CCA, respectively. These methods show drastic improvements in performance in face recognition (Zhang et al., 2006) and image-text embedding (Oshikiri et al., 2016; Fukui et al., 2016), and to otherwise.

\(\sum_{i,j} A_{ij}\) is maximized under a variance constraint.

Inspired by these existing studies, for \(D\)-view feature learning \((D \geq 1)\), PMvGE transforms data vectors \(x_i \in \mathbb{R}^{p_i}\) from view \(d_i \in \{1, \ldots, D\}\) by neural networks \(x_i \mapsto y_i := \text{NN}(d_i)(x_i) \in \mathbb{R}^{K}\) so that the function \(\exp(y_i, y_j)\) approximates the weight \(w_{ij}\) for \(i,j = 1, \ldots, n\). We introduce a parametric model of the conditional probability of \(\{w_{ij}\}\) given \(\{(x_i, d_i)\}\), and thus PMvGE is a non-linear and probabilistic extension of CDMCA. This leads to very efficient computation of the Maximum Likelihood Estimator (MLE) with minibatch SGD.

Our contribution in this paper is summarized as follows:

1. We propose PMvGE for multi-view feature learning with many-to-many associations, which is non-linear, efficient to compute, and inductive. Comparison with existing methods is shown in Table 1. See Section 2 for the description of these methods.

2. We show in Section 3 that PMvGE generalizes various existing multi-view methods, at least approximately, by considering the Maximum Likelihood Estimator (MLE) with a novel probabilistic model.

3. We show in Section 4 that PMvGE with large-scale datasets can be efficiently computed by minibatch SGD.

4. We prove that PMvGE, yet very simple, learns a wide class of similarity measures across views. By combining Mercer’s theorem and the universal approximation theorem, we prove in Section 5.1 that the inner product of feature vectors can approximate arbitrary continuous positive-definite similarity measures via sufficiently large neural networks. We also prove in Section 5.2 that MLE will actually learn the correct parameter value for sufficiently large number of
2. Related works

0-view feature learning: There are several graph embedding methods related to PMvGE without data vectors. We call them as 0-view feature learning methods. Given a graph, Spectral Graph Embedding (Chung, 1997; Belkin and Niyogi, 2001, SGE) obtains feature vectors of nodes by considering the adjacency matrix. However, SGE requires time-consuming eigenvector computation due to its variance constraint. LINE (Tang et al., 2015) is very similar to 0-view PMvGE, which reduces the time complexity of SGE by proposing a probabilistic model so that any constraint is not required. DeepWalk (Perozzi et al., 2014) and node2vec (Grover and Leskovec, 2016) obtain feature vectors of nodes by applying skip-gram model (Mikolov et al., 2013) to the node-series computed by random-walk over the given graph, while PMvGE directly considers the likelihood function.

1-view feature learning: Locality Preserving Projections (He and Niyogi, 2004; Yan et al., 2007, LPP) incorporates linear transformation of data vectors into SGE. For introducing non-linear transformation, the graph neural network (Scarselli et al., 2009, GNN) defines a graph-based convolutional neural network. ChebNet (Defferrard et al., 2016) and Graph Convolutional Network (GCN) (Kipf and Welling, 2017) reduce the time complexity of GNN by approximating the convolution. These GNN-based approaches are highly expressive but not inductive. A method is called inductive if it computes feature vectors for newly obtained vectors which are not included in the training set. Graph-SAGE (Hamilton et al., 2017) and Inductive DeepWalk (Dai et al., 2018, IDW) are inductive as well as our proposal PMvGE, but probabilistic models of these methods are different.

Multi-view feature learning: HIMFAC (Nori et al., 2012) is mathematically equivalent to CDMCA. Factorization Machine (Rendle, 2010, FM) incorporates higher-order products of multi-view data vectors to linear-regression. It can be used for link prediction across views. If only the second terms in FM are considered, FM is approximately the same as PMvGE with linear transformations. However, FM does not include PMvGE with neural networks.

Another study Stochastic Block Model (Holland et al., 1983; Nowicki and Snijders, 2001, SBM) is a well-known probabilistic model of graphs, whose links are generated with probabilities depending on the cluster memberships of nodes. SBM assumes the cluster structure of nodes, while our model does not.

3. Proposed model and its parameter estimation

3.1. Preliminaries

We consider an undirected graph consisting of $n$ nodes $\{v_i\}_{i=1}^n$ and link weights $w_{ij} \geq 0$ ($i, j = 1, 2, \ldots, n$) satisfying $w_{ij} = w_{ji}$ for all $i, j$, and $w_{ii} = 0$. Let $D \in \mathbb{N}$ be the number of views. For $D$-view feature learning, node $v_i$ belongs to one of views, which we denote as $d_i \in \{1, 2, \ldots, D\}$. The data vector representing the attributes (or side-information) at node $v_i$ is denoted as $x_i \in \mathbb{R}^{p_d}$, for view $d_i$ with dimension $p_d$. For 0-view feature learning, we formally let $D = 1$ and use the 1-hot vector $x_i \in \{0, 1\}^n$. We assume that we obtain $\{w_{ij}\}_{i,j=1}^n, \{x_i, d_i\}_{i=1}^n$ as observations. By taking $w_{ij}$ as a random variable, we consider a parametric model of conditional probability of $w_{ij}$ given the data vectors. In Section 3.2, we consider the probability model of $w_{ij}$ with the conditional expected value

$$
\mu_{ij} = E(w_{ij}|\{x_i, d_i\}_{i=1}^n)
$$

where $\{x_i, d_i\}_{i=1}^n$ is given, for all $1 \leq i < j \leq n$. In Section 3.3, we then define PMvGE by specifying the functional form of $\mu_{ij}$ via feature vectors.

3.2. Probabilistic model

For deriving our probabilistic model, we first consider a random graph model with fixed $n$ nodes. At each time-point $t = 1, 2, \ldots, T$, an unordered node pair $(v_i, v_j)$ is chosen randomly with probability

$$
\mathbb{P}(e_t = (v_i, v_j)) = \frac{\mu_{i'j'}}{\sum_{1 \leq i < j \leq n} \mu_{ij}}
$$

where $i' := \min\{i, j\}, j' := \max\{i, j\}$ and $e_t$ represents the undirected link at time $t$. The parameters $\mu_{ij} \geq 0$ ($1 \leq i < j \leq n$) are interpreted as unnormalized probabilities of node pairs. We allow the same pair is sampled several times. Given independent observations $e_1, e_2, \ldots, e_T$, we consider the number of links generated between $v_i$ and $v_j$ as

$$
w_{ij} = |\{t \in \{1, \ldots, T\} | e_t = (v_i, v_j)\}|
$$

The conditional probability $\mathbb{P}(\{w_{ij}\}_{1 \leq i < j \leq n} | T)$ follows a multinomial distribution. Assuming that $T$ obeys Poisson distribution with mean $\sum_{1 \leq i < j \leq n} \mu_{ij}$, the probability of $\{w_{ij}\}_{1 \leq i < j \leq n}$ follows

$$
\mathbb{P}(\{w_{ij}\}_{1 \leq i < j \leq n}) = \prod_{1 \leq i < j \leq n} p(w_{ij}; \mu_{ij})
$$

where $p(w; \mu)$ is the probability function of Poisson distribution with mean $\mu$. Thus $w_{ij}$ follows Poisson distribution independently as

$$
w_{ij} \sim \text{Po}(\mu_{ij})
$$

(1)
for all $1 \leq i < j \leq n$. Although $w_{ij}$ should be a non-negative integer as an outcome of Poisson distribution, our likelihood computation allows $w_{ij}$ to take any nonnegative real value.

Our probabilistic model (1) is nothing but Stochastic Block Model (Holland et al., 1983, SBM) by assuming that node $v_i$ belongs to a cluster $c_i \in \{1, 2, \ldots, C\}$. The model is specified as

$$
\mu_{ij} = \beta(c_i, c_j),
$$

(2)

where $\beta(c_i, c_j)$ is the parameter regulating the number of links whose end-points belong to clusters $c_i$ and $c_j$. Note that SBM is interpreted as 1-view method with 1-hot vector $x_i \in \{0, 1\}^C$ indicating the cluster membership.

3.3. Proposed model (PMvGE)

Inspired by various existing methods for 0-view and 1-view feature learning, we propose a novel model for the parameter $\mu_{ij}$ in eq. (1) by using the inner-product similarity as

$$
\mu_{ij}(\alpha, \psi) := \alpha^{d(d, d)} \exp \left( \langle y_i, y_j \rangle \right),
$$

(3)

$$
y_i := f^{(d)}(x_i),
$$

where $\alpha^{(d, d)} \in \mathbb{R}_{\geq 0}^{D \times D}$ is a symmetric parameter matrix ($\alpha = \alpha^\top$) for regulating the sparseness of $W = (w_{ij})$. For $y, y' \in \mathbb{R}^K$, $\langle y, y' \rangle' = \sum_{i=1}^K y_i y_i'$ is simply the inner product in Euclidean space. The functions $f^{(d)}(\cdot) : \mathbb{R}^d \rightarrow \mathbb{R}^K$, $d = 1, 2, \ldots, D$, specify the non-linear transformations from data vectors to feature vectors. $\psi$ represents a collection of parameters (e.g., neural network weights).

3.4. Link weights across some view pairs may be missing

Link weights across all the view pairs may not be available in practice. So we consider the set of unordered view pairs

$$
\mathcal{D} := \{(d, e) : \text{Link weights are observed between views } d, e\},
$$

and we formally set $w_{ij} = 0$ for the missing $(d_i, d_j) \not\in \mathcal{D}$. For example, $\mathcal{D} = \{(1, 2)\}$ for $D = 2$ indicates that link weights across view-1 and view-2 are observed while link weights within view-1 or view-2 are missing. We should notice the distinction between setting $w_{ij} = 0$ with missing and observing $w_{ij} = 0$ without missing, because these two cases give different likelihood functions.

3.5. Maximum Likelihood Estimator

Since PMvGE is specified by (1) and (3), the log-likelihood function is given by

$$
\ell_n(\alpha, \psi) := \sum_{(i, j) \in \mathcal{I}_n} \left[ w_{ij} \log \mu_{ij}(\alpha, \psi) - \mu_{ij}(\alpha, \psi) \right],
$$

(6)
whose sum is over the set of index pairs \( \mathcal{I}_n := \{(i, j) \mid 1 \leq i < j \leq n, (d_i, d_j) \in \mathcal{D}\) The Maximum Likelihood Estimator (MLE) of the parameter \((\alpha, \psi)\) is defined as the maximizer of (6). We estimate \((\alpha, \psi)\) by maximizing \(\ell_n(\alpha, \psi)\) with constraints \(\alpha = \alpha^T\) and \(\alpha(d, e) = 0\) for \((d, e) \notin \mathcal{D}\).

### 3.6. PMvGE approximately generalizes various methods for multi-view learning

SBM (2) is 1-view PMvGE with 1-hot cluster membership vector \(x_i \in \{0, 1\}^C\). Consider the linear model (5) with \(\psi^{(1)} = (\psi^1, \ldots, \psi^C) \in \mathbb{R}^{K \times C}\), where \(\psi^c \in \mathbb{R}^K\) is a feature vector for class \(c = 1, \ldots, C\). Then \(\mu_{ij} = \alpha^{(1)} \exp(\langle \psi^i, \psi^j \rangle)\) will specify \(\beta^{(c_i, c_j)}\) for sufficiently large \(K\), and \(\psi^{(c_i)}\) represent low-dimensional structure of SBM for smaller \(K\). SBM is also interpreted as PMvGE with \(D = C\) views by letting \(d_i = c_i\) and \(f^\psi(d_i)(x) \equiv 0\). Then \(\mu_{ij} = \alpha^{(c_i, c_j)}\) is equivalent to SBM.

More generally, CDMCA (Shimodaira, 2016) is approximated by \(D\)-view PMvGE with the linear transformations (5). The first half of the objective function (6) becomes

\[
\frac{1}{2} \sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij} \langle \psi^{(d_i)}(x_i), \psi^{(d_j)}(x_j) \rangle
\]

by specifying \(\alpha(d, e) \equiv 1\). CDMCA computes the transformation matrices \(\{\psi^{(d)}\}\) by maximizing this objective function under a quadratic constraint such as

\[
\sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij} \psi^{(d_i)}(x_i) \psi^{(d_j)}(x_j) = I.
\]

The above observation intuitively explains why PMvGE approximates CDMCA; this is more formally discussed in Supplement C. The quadratic constraint is required for preventing the maximizing of (7) from being diverged in CDMCA. However, our likelihood approach does not require it, because the last half of (6) serves as a regularization term.

### 3.7. PMvGE represents neural network classifiers of multiple classes

For illustrating the generality of PMvGE, here we consider the multi-class classification problem. We show that PMvGE includes Feed-Forward Neural Network (FFNN) classifier as a special case.

Let \((x_i, c_i) \in \mathbb{R}^d \times \{1, \ldots, C\}\), \(i = 1, \ldots, n\) be the training data for the classification problem of \(C\) classes. FFNN classifier with softmax function (Bishop, 2006) is defined as

\[
h_j(x_i) := \frac{\exp(f_{\psi,j}(x_i))}{\sum_{j=1}^{C} \exp(f_{\psi,j}(x_i))}, \quad j = 1, \ldots, C,
\]

where \(f_{\psi}(x) := (f_{\psi,1}(x), \ldots, f_{\psi,C}(x)) \in \mathbb{R}^{C}\) is a multi-valued neural network, and \(x_i\) is classified into the class \(\arg \max_{j \in \{1, \ldots, C\}} h_j(x_i)\). This classifier is equivalent to

\[
\arg \max_{j \in \{1, \ldots, C\}} \exp (\langle f_{\psi}(x_i), e_j \rangle),
\]

where \(e_j \in \{0, 1\}^C\) is the 1-hot vector with 1 at \(j\)-th entry and 0 otherwise.

The classifier (9) can be interpreted as PMvGE with \(D = 2\), \(\mathcal{D} = \{(1, 2)\}\) as follows. For view-1, \(f_{\psi}^{(1)}(x) := f_{\psi}(x)\) and inputs are \(x_1, \ldots, x_n\). For view-2, \(f_{\psi}^{(2)}(x') := x'\) and inputs are \(e_1, \ldots, e_C\). We set \(w_{ij} = 1\) between \(x_i\) and \(e_{c_i}\), and \(w_{ij} = 0\) otherwise.

### 4. Optimization

In this section, we present an efficient way of optimizing the parameters for maximizing the objective function \(\ell_n(\alpha, \psi)\) defined in eq. (6). We alternatively optimize the two parameters \(\alpha\) and \(\psi\). Efficient update of \(\psi\) with minibatch SGD is considered in Section 4.1, and update of \(\alpha\) by solving an estimating equation is considered in Section 4.2. We iterate these two steps for maximizing \(\ell_n(\alpha, \psi)\).

#### 4.1. Update of \(\psi\)

We update \(\psi\) using the gradient of \(\ell_n(\alpha, \psi)\) by fixing current parameter value \(\alpha\). Since \(W = (w_{ij})\) may be sparse in practice, the computational cost of minibatch SGD is expected to be reduced by considering the sum over the set \(\mathcal{W}_n := \{(i, j) \in \mathcal{I}_n \mid w_{ij} > 0\}\). On the other hand, there should be \(|\mathcal{I}_n| = O(n^2)\) positive terms in the last half of \(\ell_n(\alpha, \psi)\), so we consider the sum over node pairs uniformly-resampled from \(\mathcal{I}_n\).

We make two sets \(\mathcal{I}_n', \mathcal{W}_n'\) by picking \((i, j)\) from \(\mathcal{I}_n\) and \(\mathcal{W}_n\), respectively, so that

\[
|\mathcal{I}_n'| + |\mathcal{W}_n'| = m, \quad |\mathcal{I}_n'|/|\mathcal{W}_n'| = r.
\]

User-specified constants \(m \in \mathbb{N}\) and \(r > 0\) are usually called as “minibatch size” and “negative sampling rate”. We sequentially update minibatch \(\mathcal{I}_n', \mathcal{W}_n'\) for computing the gradient of

\[
\sum_{(i, j) \in \mathcal{W}_n'} w_{ij} \log \mu_{ij}(\alpha, \psi) - \tau \sum_{(i, j) \in \mathcal{I}_n'} \mu_{ij}(\alpha, \psi)
\]

with respect to \(\psi\). By utilizing the gradient, the parameter \(\psi\) can be sequentially updated by SGD, where \(\tau > 0\) is a tuning parameter. Eq. (10) approximates \(\ell_n(\alpha, \psi)\) if \((i, j)\) are uniformly-resampled and \(\tau = |\mathcal{I}_n'|/|\mathcal{W}_n'|\) however, smaller \(\tau\) such as \(\tau = 1\) may make this algorithm stable in some cases.
4.2. Update of $\alpha$

Let $\phi$ represent current parameter value of $\psi$. By solving the estimating equation $\frac{\partial \ell_n(\psi, \phi)}{\partial \alpha} = 0$ with respect to $\alpha$ under constraints $\alpha = \alpha^\dagger$ and $\alpha^\dagger = 0$, $(d, e) \notin D$, we explicitly obtain a local maximizer of $\ell_n(\alpha, \psi)$. However, the local maximizer requires roughly $O(n^2)$ operations for computation. To reduce the high computational cost, we efficiently update $\alpha$ by (11), which is a minibatch-based approximation of the local maximizer.

$$\hat{\alpha}^{(d,e)} := \begin{cases} \frac{\sum_{(i,j) \in \mathcal{I}_n^{(d,e)}} w_{ij}}{\sum_{(i,j) \in \mathcal{I}_n^{(d,e)}} \exp(\bar{g}_{ij})} & \text{for } (d,e) \in D \\ 0 & \text{otherwise} \end{cases} \tag{11}$$

where $\bar{g}_{ij} := \langle f^{(d)}(x_i), f^{(e)}(x_j) \rangle$, $\mathcal{I}_n^{(d,e)} := \{(i,j) \in \mathcal{I}_n \mid (d,j) = (d,e)\}$, and $\mathcal{I}_n$ is the minibatch defined in Section 4.1. Note that $(d,e)$ is unordered view pair.

4.3. Computational cost

PMvGE requires $O(m)$ operations for each minibatch iteration. It is efficiently computed even if the number of data vectors $n$ is very large.

5. PMvGE learns arbitrary similarity measure

Two theoretical results are shown here for indicating that PMvGE with sufficiently large neural networks learns arbitrary similarity measure using sufficiently many data vectors. In Section 5.1, we prove that arbitrary similarity measure can be approximated by the inner product in the shared space with sufficiently large neural networks. In Section 5.2, we prove that MLE of PMvGE converges to the true parameter value, i.e., the consistency of MLE, in some sense as the number of data vectors increases.

5.1. Inner product of NNs approximates a wide class of similarity measures across views

Feedforward neural networks with ReLU or sigmoid functions are proved to be able to approximate arbitrary continuous functions under some assumptions (Cybenko, 1989; Funahashi, 1989; Yarotsky, 2016; Telgarsky, 2017). However, these results cannot be directly applied to PMvGE, because our model is based on the inner product of two neural networks $\langle f^{(d)}(x_i), f^{(d)}(x_j) \rangle$. In Theorem 5.1, we show that the inner product can approximate $g_\ast \left( f^{(d)}(x_i), f^{(d)}(x_j) \right)$, i.e., arbitrary similarity measure $g_\ast(\cdot, \cdot)$ in $K_\ast$ dimensional shared space, where $f^{(d)}(x_i), f^{(d)}(x_j)$ are arbitrary two continuous functions. For showing $g_\ast \left( f^{(d)}(x), f^{(e)}(x') \right) \approx \langle f^{(d)}(x), f^{(e)}(x') \rangle$, the idea is first consider a feature map $\Phi_K : \mathbb{R}^{K} \rightarrow \mathbb{R}^{K}$ for approximating $g_\ast(y_s, y'_s) \approx \langle \Phi_K(y_s), \Phi_K(y'_s) \rangle$ with sufficiently large $K$, and then consider neural networks $f^{(d)}_\psi : \mathbb{R}^{p_d} \rightarrow \mathbb{R}^{K}$ with sufficiently many hidden units for approximating $\Phi_K(f^{(d)}(x)) \approx f^{(d)}(x)$.

**Theorem 5.1** Let $f^{(d)}_\psi : [-M, M]^{p_d} \rightarrow [-M', M']^{K}$, $d = 1, 2, \ldots, D$, be continuous functions and $g_\ast : [-M', M']^{K} \times [-M', M']^{K} \rightarrow \mathbb{R}$ be a symmetric, continuous, and positive-definite kernel function for some $K^*, M, M' > 0$. $\sigma(\cdot)$ is ReLU or activation function which is non-constant, continuous, bounded, and monotonically-increasing. Then, for arbitrary $\varepsilon > 0$, by specifying sufficiently large $K \in \mathbb{N}, T = T(K) \in \mathbb{N}$, there exist $A^d \in \mathbb{R}^{K \times T}, B^d \in \mathbb{R}^{T \times p_d}, c \in \mathbb{R}^T, d \in \{1, 2, \ldots, D\}$, such that

$$\left| g_\ast \left( f^{(d)}_\psi(x), f^{(e)}_\psi(x') \right) - \langle f^{(d)}(x), f^{(e)}(x') \rangle \right| < \varepsilon \tag{12}$$

for all $(x, x') \in [-M, M]^{p_d + p_e}, d, e \in \{1, 2, \ldots, D\}$, where $f^{(d)}_\psi(x^d) = A^d \sigma(B^d x^d + c^d)$, $d = 1, 2, \ldots, D$, are two-layer neural networks with $T$ hidden units and $\sigma(x)$ is element-wise $\sigma(\cdot)$ function.

Proof of the theorem is given in Supplement A.

If $D = 1$, Theorem 5.1 corresponds to Mercer’s theorem (Mercer, 1909; Courant and Hilbert, 1989) of Kernel methods, which states that arbitrary positive definite kernel $g_\ast(\cdot, \cdot)$ can be expressed as the inner product of high-dimensional feature maps. While Mercer’s theorem indicates only the existence of such feature maps, Theorem 5.1 also states that the feature maps can be approximated by neural networks.

**Illustrative example** As an example of positive definite similarity, we consider the cosine similarity

$$g_\ast(f_s(x), f_s(x')) := \frac{\langle f_s(x), f_s(x') \rangle}{\|f_s(x)\|_2 \|f_s(x')\|_2}$$

with $f_s(x) = (x_1, \cos x_2, \exp(-x_3), \sin(x_4 - x_5), p = 5$. For 2-dim visualization in Fig. 2 with $(s, t) \in \mathbb{R}^2$, let us define $G_s(s, t) := g_\ast(f_s(se_1), f_s(te_2)), e_1 := (1, 1, 0, 0), e_2 := (0, 0, 1, 1, 1)$ and its approximation by the inner product of neural networks $\hat{G}_K(s, t) := \langle \phi(s e_1), \phi(t e_2) \rangle$ with $T = 10^3$ hidden units. If $K$ and $T$ are sufficiently large, $\hat{G}_K(s, t)$ approximates $G_s(s, t)$ very well as suggested by Theorem 5.1.

5.2. MLE converges to the true parameter value

We have shown the universal approximation theorem of similarity measure in Theorem 5.1. However, it only states that the good approximation is achieved if we properly tune the parameters of neural networks. Here we argue that MLE
of Section 3.5 will actually achieve the good approximation if we have sufficiently many data vectors. The technical details of the argument are given in Supplement B.

Let $\theta \in \Theta$ denote the vector of free parameters in $\alpha, \psi$, and $\ell_n(\theta)$ be the log-likelihood function (6). We assume that the optimization algorithm in Section 4 successfully computes MLE $\hat{\theta}_n$ that maximizes $\ell_n(\theta)$. Here we ignore the difficulty of global optimization, while we may only get a local maximum in practice. We also assume that PMvGE is correctly specified; there exists a parameter value $\theta$ so that the parametric model represents the true probability distribution.

Then, we would like to claim that $\hat{\theta}_n$ converges to the true parameter value $\theta_*$ in the limit of $n \to \infty$, the property called the consistency of MLE. However, we have to pay careful attention to the fact that PMvGE is not a standard setting in the sense that (i) there are correlated $O(n^2)$ samples instead of $n$ i.i.d. observations, and (ii) the model is not identifiable with infinitely many equivalent parameter values; for example there are rotational degrees of freedom in the shared space so that $(\mathbf{y}, \mathbf{y}') = (\mathbf{Oy}, \mathbf{Oy}')$ with any orthogonal matrix $\mathbf{O}$ in $\mathbb{R}^K$. We then consider the set of equivalent parameters $\hat{\Theta}_n := \{ \theta \in \Theta \mid \ell_n(\theta) = \ell_n(\hat{\theta}_n) \}$. Theorem B.2 states that, as $n \to \infty$, $\hat{\Theta}_n$ converges to $\Theta_*$, the set of $\theta$ values equivalent to $\theta_*$. 

6. Real data analysis

6.1. Experiments on Citation dataset (1-view)

Dataset: We use Cora dataset (Sen et al., 2008) of citation network with 2,708 nodes and 5,278 ordered edges. Each node $v_i$ represents a document, which has 1,433-dimensional (bag-of-words) data vector $x_i \in \{0, 1\}^{1433}$ and a class label of 7 classes. Each directed edge represents citation from a document $v_i$ to another document $v_j$. We set the link weight as $w_{ij} = w_{ji} = 1$ by ignoring the direction, and $w_{ij} = 0$ otherwise. There is no cross or self-citation. We divide the dataset into training set consisting of 2,166 nodes (80%) with their edges, and test set consisting of remaining 542 nodes (20%) with their edges. We utilize 20% of the training set for validation. Hyper-parameters are tuned by utilizing the validation set.

We compare PMvGE with several feature learning methods: Stochastic Block Model (Holland et al., 1983, SBM), ISOMAP (Tenenbaum et al., 2000), Locally Linear Embedding (Roweis and Saul, 2000, LLE), Spectral Graph Embedding (Belkin and Niyogi, 2001, SGE), Multi Dimensional Scaling (Kruskal, 1964, MDS), DeepWalk (Perozzi et al., 2014), and GraphSAGE (Hamilton et al., 2017).

NN for PMvGE: 2-layer fully-connected network, which consists of 3,000 tanh hidden units and 1,000 tanh output units, is used. The network is trained by Adam (Kingma and Ba, 2015) with batch normalization. The learning rate is starting from 0.0001 and attenuated by 1/10 for every 100 iterations. Negative sampling rate $r$ and minibatch size $m$ are set as 1 and 512, respectively, and the number of iterations is 200.

Parameter tuning: For each method, parameters are tuned on validation sets. Especially, the dimension of feature vectors is selected from {50, 100, 150, 200}.

Label classification (Task 1): We classify the documents into 7 classes using logistic regression with the feature vector as input and the class label as output. We utilize LibLinear (Fan et al., 2008) for the implementation.

Clustering (Task 2): The $k$-means clustering of the feature vectors is performed for unsupervised learning of document clusters. The number of clusters is set as 7.

Results: The quality of classification is evaluated by classification accuracy in Task 1, and Normalized Mutual Information (NMI) in Task 2. Sample averages and standard deviations over 10 experiments are shown in Table 2. In experiment (A), we apply methods to both training set and test set, and evaluate them by test set. In (B), we apply methods to only the training set, and evaluate them by test set. SGE, MDS, and DeepWalk are not inductive, and they cannot be applied to unseen data vectors in (B). PMvGE outperforms the other methods in both experiments.

6.2. Experiments on AwA dataset (2-view)

Dataset: We use Animal with Attributes (AwA) dataset (Lampert et al., 2009) with 30,475 images for view-1 and 85 attributes for view-2. We prepared 4,096 dimensional DeCAF data vector (Donahue et al., 2014) for each image, and 300 dimensional GloVe (Pennington et al., 2014) data vector for each attribute. Each image is associated with some attributes. We set $w_{ij} = 1$ for the associated pairs.
We compare PMvGE with CCA, DCCA (Andrew et al., 2013), SGE, DeepWalk, and GraphSAGE.

**NN for PMvGE:** Each view uses a 2-layer fully-connected network, which consists of 2000 tanh hidden units and 100 tanh output units. The dimension of the feature vector is $K = 100$. Adam is used for optimization with Batch normalization and Dropout ($p = 0.5$). Minibatch size, learning rate, and momentum are tuned on the validation set. We monitor the score on the validation set for early stopping.

**Parameter tuning:** For each method, parameters are tuned on validation sets. Especially, the dimension of feature vectors is selected from $\{10, 50, 100, 150\}$.

**Link prediction (Task 3):** For each query image, we rank attributes according to the cosine similarity of feature vectors across views. An attribute is regarded as correct if it is associated with the query image.

**Results:** The quality of the ranked list of attributes is measured by Average Precision (AP) score in Task 3. Sample averages and standard deviations over 10 experiments are shown in Table 2. In experiment (A), we apply methods to both training set and test set, and evaluate them by test set. The whole training set is used for validation. In experiment (B), we apply methods to only the training set, and evaluate them by test set. 20% of training set is used for validation. PMvGE outperforms the other methods including DCCA. While DeepWalk shows good performance in experiment (A), DeepWalk and SGE cannot be applied to unseen data vectors in (B). Unlike SGE and DeepWalk which only consider the associations, 1-view feature learning methods such as GraphSAGE cannot be applied to this AwA dataset since the dimension of data vectors is different depending on the view. So we do not perform 1-view methods in Task 3.

**Locality of each-view is preserved through neural networks:** To see whether the locality of input is preserved through neural networks in PMvGE, we computed the Spearman’s rank correlation coefficient between $(x, x')$ and $(f_w^{(d)}(x), f_w^{(d)}(x'))$ for view-d data vectors $x, x'$ in AwA dataset ($d = 1, 2$). For DeCAF (view-1) and GloVe (view-2) inputs, the values are $0.722 \pm 0.058$ and $0.811 \pm 0.082$, respectively. This result indicates that the feature vectors of PMvGE preserves the similarities of the data vectors fairly well.

### Table 2: Task 1 and Task 2 for the experiment on Citation dataset ($D = 1$), and Task 3 for the experiment on AwA dataset ($D = 2$). The larger values are better.

|                  | (A)             | (B)             |
|------------------|-----------------|-----------------|
| **Task 1** ($D = 1$) |                 |                 |
| ISOMAP           | 54.5 ± 1.78     | 54.8 ± 2.43     |
| LLE              | 30.2 ± 1.91     | 31.9 ± 2.62     |
| SGE              | 47.6 ± 1.64     | -               |
| MDS              | 29.8 ± 2.25     | -               |
| DeepWalk         | 54.2 ± 2.04     | -               |
| GraphSAGE        | 60.8 ± 1.73     | 57.1 ± 1.61     |
| **PMvGE**        | **74.8 ± 2.55** | **71.1 ± 2.10** |
| **Task 2** ($D = 1$) |                 |                 |
| SBM              | 4.37 ± 1.44     | 2.81 ± 0.10     |
| ISOMAP           | 13.0 ± 0.36     | 14.3 ± 1.98     |
| LLE              | 7.40 ± 3.40     | 9.47 ± 3.00     |
| SGE              | 1.41 ± 0.34     | -               |
| MDS              | 2.81 ± 0.10     | -               |
| DeepWalk         | 16.7 ± 1.05     | -               |
| GraphSAGE        | 19.6 ± 0.93     | 12.4 ± 3.00     |
| **PMvGE**        | **35.9 ± 0.88** | **30.5 ± 3.90** |
| **Task 3** ($D = 2$) |                 |                 |
| CCA              | 45.5 ± 0.20     | 42.4 ± 0.30     |
| DCCA             | 41.4 ± 0.30     | 41.2 ± 0.35     |
| SGE              | 43.5 ± 0.39     | -               |
| DeepWalk         | 71.3 ± 0.57     | -               |
| **PMvGE**        | **71.5 ± 0.48** | **70.5 ± 0.53** |

### 7. Conclusion

We presented a simple probabilistic framework for multi-view learning with many-to-many associations. We name the framework as Probabilistic Multi-view Graph Embedding (PMvGE). Various existing methods are approximately included in PMvGE. We gave theoretical justification and practical estimation algorithm to PMvGE. Experiments on real-world datasets showed that PMvGE outperforms existing methods.

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Supplementary Material: A probabilistic framework for multi-view feature learning with many-to-many associations via neural networks

A. Proof of Theorem 5.1

Since \( g_\ast : [-M', M']^{2K} \to \mathbb{R} \) is a positive definite kernel on a compact set, it follows from Mercer’s theorem that there exist positive eigenvalues \( \{ \lambda_k \}_{k=1}^\infty \) and continuous eigenfunctions \( \{ \phi_k \}_{k=1}^\infty \) such that

\[
g_\ast(y_\ast, y'_\ast) = \sum_{k=1}^\infty \lambda_k \phi_k(y_\ast) \phi_k(y'_\ast), \quad y_\ast, y'_\ast \in [-M', M']^{K'},
\]

where the convergence is absolute and uniform (Minh et al., 2006). The uniform convergence implies that for any \( \varepsilon_1 > 0 \) there exists \( K_0 \in \mathbb{N} \) such that

\[
\sup_{(y_\ast, y'_\ast) \in [-M', M']^{2K}} \left| g_\ast(y_\ast, y'_\ast) - \sum_{k=1}^K \lambda_k \phi_k(y_\ast) \phi_k(y'_\ast) \right| < \varepsilon_1, \quad K \geq K_0.
\]

This means \( g_\ast(y_\ast, y'_\ast) \approx (\Phi_K(y_\ast), \Phi_K(y'_\ast)) \) for a feature map \( \Phi_K(y_\ast) = (\sqrt{\lambda_k} \phi_k(y_\ast))_{k=1}^K \).

We fix \( K \) and consider approximation of \( h_k^{(d)}(x) = \sqrt{\lambda_k} \phi_k(f_k^{(d)}(x)) \) below. Since \( h_k^{(d)} \) are continuous functions on a compact set, there exists \( C = C(K) > 0 \) such that

\[
\sup_{x \in [-M, M]^d} |h_k^{(d)}(x)| < C, \quad k = 1, \ldots, K, d = 1, \ldots, D.
\]

Let us write the neural networks as \( f_k^{(d)} = (f_1^{(d)}, \ldots, f_d^{(d)}) \), where \( f_k^{(d)} : \mathbb{R}^d \to \mathbb{R}, \ d = 1, \ldots, D, \ k = 1, \ldots, K, \) are two-layer neural networks with \( T \) hidden units. Since \( h_k^{(d)} \) are continuous functions, it follows from the universal approximation theorem (Cybenko, 1989; Telgarsky, 2017) that for any \( \varepsilon_2 > 0 \), there exists \( T_0(K) \in \mathbb{N} \) such that

\[
\sup_{x \in [-M, M]^d} \left| h_k^{(d)}(x) - f_k^{(d)}(x) \right| < \varepsilon_2, \quad k = 1, \ldots, K, \ d = 1, \ldots, D
\]

for \( T \geq T_0(K) \). Therefore, for all \( d, e \in \{1, 2, \ldots, D\} \), we have

\[
\sup_{(x, x') \in [-M, M]^d \times [-M, M]^d} \left| g_\ast\left(f_\ast^{(d)}(x), f_\ast^{(e)}(x')\right) - \sum_{k=1}^K h_k^{(d)}(x) h_k^{(e)}(x') - \sum_{k=1}^K f_k^{(d)}(x) f_k^{(e)}(x') \right|
\]

\[
\leq \sup_{(x, x') \in [-M, M]^d \times [-M, M]^d} \left| g_\ast\left(f_\ast^{(d)}(x), f_\ast^{(e)}(x')\right) - \sum_{k=1}^K h_k^{(d)}(x) h_k^{(e)}(x') - \sum_{k=1}^K f_k^{(d)}(x) f_k^{(e)}(x') \right|
\]

\[
+ \sup_{(x, x') \in [-M, M]^d \times [-M, M]^d} \left| \sum_{k=1}^K h_k^{(d)}(x) (h_k^{(e)}(x') - f_k^{(e)}(x')) \right|
\]

\[
+ \sup_{(x, x') \in [-M, M]^d \times [-M, M]^d} \left| \sum_{k=1}^K (h_k^{(d)}(x) - f_k^{(d)}(x)) f_k^{(e)}(x') \right|
\]

\[
\leq \sup_{y_\ast, y'_\ast \in [-M', M']^{K'}} \left| g_\ast(y_\ast, y'_\ast) - \sum_{k=1}^K \lambda_k \phi_k(y_\ast) \phi_k(y'_\ast) \right|
\]

\[
+ \sup_{x \in [-M, M]^d} \left| h_k^{(d)}(x) \right| \sup_{x' \in [-M, M]^d} \left| h_k^{(e)}(x') - f_k^{(e)}(x') \right|
\]

\[
+ \sup_{x \in [-M, M]^d} \left| h_k^{(d)}(x) - f_k^{(d)}(x) \right| \sup_{x' \in [-M, M]^d} \left| f_k^{(e)}(x') \right|
\]

\[
< \varepsilon_1 + KC\varepsilon_2 + K\varepsilon_2(C + \varepsilon_2).
\]
By letting $\varepsilon_1 = \varepsilon/2, \varepsilon_2 = \min(C, \varepsilon/(6K C))$, the last formula becomes smaller than $\varepsilon$, thus proving

$$\sup_{(x,x') \in [-M,M]^{d + p + r_d}} \left| g_s \left( f_s^{(d)}(x), f_s^{(e)}(x') \right) - \sum_{k=1}^{K} f_k^{(d)}(x) f_k^{(e)}(x') \right| < \varepsilon, \quad d, e = 1, \ldots, D. $$

The two conditions (i) and (iii) hold obviously, and thus we verify (ii) and (iv) below.

### B. Consistency of MLE in PMvGE

In this section, we provide technical details of the argument of Section 5.2.

For proving the consistency of MLE, we introduce the following generative model. Let $d_i, i = 1, \ldots, n$, be random variables independently distributed with the probability $P(d_i = d) = \eta^{(d)} \in (0, 1)$ where $\sum_{d=1}^{D} \eta^{(d)} = 1$. Data vectors are also treated as random variables. The conditional distribution of $x_i$ given $d_i$ is

$$x_i | d_i \overset{\text{indep.}}{\sim} q^{(d_i)}, \quad i = 1, \ldots, n,$$

where $q^{(d_i)}$ is a distribution on a compact support in $\mathbb{R}^{p_d}$. Let us denote eq. (3) as $\mu_{ij}(x_i, x_j, d_i, d_j | \alpha, \psi)$ for indicating the dependency on $(x_i, x_j, d_i, d_j)$. The conditional distributions of link weights are already specified in (1) as

$$w_{ij} | x_i, x_j, d_i, d_j \overset{\text{indep.}}{\sim} \text{Po}(\mu_{ij}^*), \quad i, j = 1, \ldots, n,$$

where $\mu_{ij}^* = \mu_{ij}(x_i, x_j, d_i, d_j | \alpha, \psi)$ with a true parameter $(\alpha, \psi)$. Due to the constraints $\alpha = \alpha^\top$ and $w_{ij} = 0 ((d_i, d_j) \notin D)$, the vector of free parameters in $\alpha$ is $\alpha_D := \{ \alpha^{(d,e)} \}_{(d,e) \in D, d \leq e} \in \mathbb{R}^{d D}$, and we write $\theta := (\alpha_D, \psi)$. Let $\tilde{n} := |I_n| = O(n^2)$ denote the number of terms in the sum of $\ell_n(\theta)$ in eq. (6). Then the expected value of $\tilde{n}^{-1} \ell_n(\theta)$ under the generative model with the true parameter $\theta_*$ is expressed as

$$\ell(\theta) := E_{x_1, x_2, \ldots, x_n} \left[ \mu_1(x_1, x_2, x_3, x_4) \log \mu_2(x_1, x_2, x_3, x_4) - \mu_3(x_1, x_2, x_3, x_4) \right].$$

If it were the case of i.i.d. observations of sample size $n$, we would have that $\tilde{n}^{-1} \ell_n(\theta)$ converges to $\ell(\theta)$ as $n \to \infty$ from the law of large numbers. In Theorem B.1, we actually prove the uniform convergence in probability, but we have to pay careful attention to the fact that $n^2$ observations of $(x_i, x_j, d_i, d_j)$ are not independent when indices overlap.

**Theorem B.1** Let us assume that the parameter space of $\theta$ is $\Theta := [\delta, 1/\delta]^{|D|} \times \Psi$, where $\delta \in (0, 1)$ is a sufficiently small constant and $\Psi \subset \mathbb{R}^q$ is a compact set. Assume also that the transformations $f^{(d)}(x)$, $d = 1, \ldots, D$, are Lipschitz continuous with respect to $(\psi, x)$. Then we have, as $n \to \infty$,

$$\sup_{\theta \in \Theta} |\tilde{n}^{-1} \ell_n(\theta) - \ell(\theta)| \overset{P}{\to} 0. \quad (13)$$

**Proof B.1** We refer to Corollary 2.2 in Newey (1991). This corollary shows $\sup_{\theta \in \Theta} |\hat{Q}_n(\theta) - \tilde{Q}_n(\theta)| = o_p(1)$ under general setting of $\hat{Q}_n(\theta)$ and $\tilde{Q}_n(\theta)$. Here we consider the case of $\hat{Q}_n(\theta) = \tilde{n}^{-1} \ell_n(\theta)$ and $\tilde{Q}_n(\theta) = \ell(\theta)$. For showing (13), the four conditions of the corollary are written as follows. (i) $\Theta$ is compact, (ii) $\tilde{n}^{-1} \ell_n(\theta) - \ell(\theta) \overset{L_p}{\to} 0$ for each $\theta \in \Theta$, (iii) $\ell(\theta)$ is continuous, (iv) there exists $B_n = O_p(1)$ such that $|\tilde{n}^{-1} \ell_n(\theta) - \tilde{n}^{-1} \ell_n(\theta')| \leq B_n \|\theta - \theta'\|_2$ for all $\theta, \theta' \in \Theta$. The two conditions (i) and (iii) hold obviously, and thus we verify (ii) and (iv) below.

Before verifying (ii) and (iv), we first consider an array $Z := (Z_{ij})$ of random variables $Z_{ij} \in Z$, $(i, j) \in I_n$, and a bounded and continuous function $h : \mathbb{R} \to \mathbb{R}$. We assume that $Z \subset \mathbb{R}$ is a compact set, and $Z_{ij}$ is independent of $Z_{kl}$ if
By considering \(|I_n \setminus R_n(i, j)| = O(n)|, the last formula is \(O(n^{-2} \cdot n) = O(n^{-1})\). Therefore,

\[
V_Z \left[ \frac{1}{n} \sum_{(i,j) \in I_n} h(Z_{ij}) \right] = O(n^{-1}). \tag{14}
\]

Next we evaluate the variance of \(\tilde{\ell}_n(\theta)\) to show (ii). Denoting \(W := (w_{ij}), X := (x_i), d := (d_i)\),

\[
V_{W,X,d}[\tilde{\ell}_n(\theta)] = E_{X,d}[V_W[\tilde{\ell}_n(\theta) \mid X, d]] + V_{X,d}[E_W[\tilde{\ell}_n(\theta) \mid X, d]]
\]

for every \(\theta \in \Theta\). The first term in the last formula is \(O(n^{-1} \cdot n) = O(n^{-1}) = O(1)\), and the second term is \(O(n^{-1}) = O(1)\) by applying eq. (14) with \(Z_{ij} := (x_i, x_j, d_i, d_j), h(Z_{ij}) = \mu_{ij}(\theta_*) \log \mu_{ij}(\theta) - \mu_{ij}(\theta)\). Therefore, \(V_{W,X,d}[\tilde{\ell}_n(\theta)] = O(1)\) and Chebyshev’s inequality implies the pointwise convergence \(\tilde{\ell}_n(\theta) \xrightarrow{p} \ell(\theta)\) for every \(\theta \in \Theta\) where \(\ell(\theta) = E_{W,X,d}[\tilde{\ell}_n(\theta)] = E_{x_1, x_2, d_1, d_2}[(\mu_{12}(\theta_*) \log \mu_{12}(\theta) - \mu_{12}(\theta))]\). Thus, condition (ii) holds.

Finally, we work on condition (iv). Since \(\mu_{ij}(\theta)\) is a composite function of \(\Theta\), \(\mu_{ij}(\theta)\) is Lipschitz continuous. The Lipschitz continuity of \(\mu_{ij}(\theta)\) and \(\mu_{ij}(\theta) > 0 (\theta \in \Theta)\) indicates the Lipschitz continuity of \(\log \mu_{ij}(\theta)\). Therefore, there exist \(M_1, M_2 > 0\) such that

\[
|\tilde{\ell}_n(\theta) - \tilde{\ell}_n(\theta')| \leq \frac{1}{n} \sum_{(i,j) \in I_n} w_{ij} (\log \mu_{ij}(\theta) - \log \mu_{ij}(\theta')) - \frac{1}{n} \sum_{(i,j) \in I_n} (\mu_{ij}(\theta) - \mu_{ij}(\theta'))
\]

\[
\leq \frac{1}{n} \sum_{(i,j) \in I_n} w_{ij} |\log \mu_{ij}(\theta) - \log \mu_{ij}(\theta')| + \frac{1}{n} \sum_{(i,j) \in I_n} |\mu_{ij}(\theta) - \mu_{ij}(\theta')|
\]

\[
\leq M_1 \left( \tilde{\ell}_n \sum_{(i,j) \in I_n} w_{ij} \right) \|\theta - \theta'\|_2 + M_2 \|\theta - \theta'\|_2.
\]

Denoting \(B_n := M_1 \cdot \tilde{\ell}_n \sum_{i,j} w_{ij} + M_2\), we have

\[
|\tilde{\ell}_n(\theta) - \tilde{\ell}_n(\theta')| \leq B_n \|\theta - \theta'\|_2.
\]

Since \(\tilde{\ell}_n \sum_{i,j} w_{ij} = O_p(1)\), the law of large numbers indicates \(B_n = O_p(1)\). Thus, condition (iv) holds. \(\square\)

Noticing that \(\theta_*\) is a maximizer of \(\ell(\theta)\) and \(\tilde{\theta}_n\) is a maximizer of \(\ell_n(\theta)\), we would have the desired result \(\tilde{\theta}_n \xrightarrow{p} \theta_*\) by combining Theorem B.1 and continuity of \(\ell(\theta)\). However it does not hold unfortunately. Instead, we define the set of parameter values equivalent to \(\theta_*\) as \(\Theta_* := \{\theta \in \Theta \mid \ell(\theta) = \ell(\theta_*)\}\). Every \(\theta \in \Theta_*\) gives the correct probability of link weights, because \(\ell(\theta) = \ell(\theta_*)\) holds if and only if \(\mu_{12}(\theta) = \mu_{12}(\theta_*)\) almost surely w.r.t. \((x_1, x_2, d_1, d_2)\). With this setting, the theorem below states that \(\tilde{\theta}_n\) converges to \(\Theta_*\) in probability. This indicates that, \(\tilde{\theta}_n\) will represent the true probability model for sufficiently large \(n\).
Theorem B.2. Let \( d_H(\cdot, \cdot) \) denote the Hausdorff distance defined as \( \max - \min L^2 \)-distance between two sets. We assume the same conditions as in Theorem B.1. Then we have, as \( n \to \infty \),
\[
d_H(\hat{\Theta}_n, \Theta_*) \xrightarrow{p} 0. \tag{15}
\]

Proof B.2. We refer to the case (1) of Theorem 3.1 in Chernozhukov et al. (2007) with \( \hat{c} = 1 \) under the condition C.1 This theorem shows, for general setting of \( \hat{\Theta}_I, \Theta_I \), that
\[
d_H(\hat{\Theta}_I, \Theta_I) \xrightarrow{p} 0. \tag{16}
\]

Here \( \hat{\Theta}_I := \{ \theta \in \Theta | \hat{Q}_n(\theta) \leq 1/a_n \} \) and \( \Theta_I := \arg \inf_{\theta \in \Theta} Q(\theta) \), where \( \hat{Q}_n(\theta), Q(\theta) \) are general functions satisfying \( \sup_{\theta \in \Theta} Q_n(\theta) = o_p(1/a_n) \), and \( a_n \to \infty \).

For proving (15), we consider the case of \( \hat{Q}_n(\theta) = -\hat{n}^{-1}\ell_n(\theta) + \sup_{\theta \in \Theta} \hat{n}^{-1}\ell_n(\theta), Q(\theta) = -\ell(\theta) + \sup_{\theta \in \Theta} \ell(\theta) \). The condition C.1 for (16) is re-written as follows. (i) \( \Theta \) is a (non-empty) compact set, (ii) \( \hat{n}^{-1}\ell_n(\theta) \) and \( \ell(\theta) \) are continuous, (iii) \( \sup_{\theta \in \Theta} \hat{n}^{-1}\ell_n(\theta) - \ell(\theta) \xrightarrow{p} 0 \), and (iv) \( \sup_{\theta \in \Theta} (-\hat{n}^{-1}\ell_n(\theta) + \sup_{\theta \in \Theta} \hat{n}^{-1}\ell_n(\theta)) \xrightarrow{p} 0 \). The conditions (i), (ii) are obvious. (iii) is shown in Theorem B.1. (iv) is verified by
\[
\sup_{\theta \in \Theta} \left( -\hat{n}^{-1}\ell_n(\theta) + \sup_{\theta \in \Theta} \hat{n}^{-1}\ell_n(\theta) \right) = -\inf_{\theta \in \Theta} \hat{n}^{-1}\ell_n(\theta) + \sup_{\theta \in \Theta} \hat{n}^{-1}\ell_n(\theta) \xrightarrow{p} -\ell(\theta_*) + \ell(\theta_*) = 0,
\]
where \( \theta_* \) is an element of \( \Theta_I \). Thus, (16) holds.

Next, we consider two sets \( \hat{\Theta}_n = \arg \sup_{\theta \in \Theta} \hat{n}^{-1}\ell_n(\theta) = \{ \theta \in \Theta | Q_n(\theta) = 0 \} \) and \( \hat{\Theta}_I \). Since these sets satisfy \( Q_n(\theta) = 0 (\theta \in \hat{\Theta}_n), Q_n(\theta') \leq 1/a_n (\theta' \in \hat{\Theta}_I) \) and \( 1/a_n \to 0 \) as \( n \to \infty \), we have \( d_H(\hat{\Theta}_n, \hat{\Theta}_I) \xrightarrow{p} 0 \). It follows from this convergence and (16) that, by noticing \( \Theta_* = \Theta_I \),
\[
d_H(\hat{\Theta}_n, \Theta_*) = d_H(\hat{\Theta}_n, \hat{\Theta}_I) \leq d_H(\hat{\Theta}_n, \hat{\Theta}_I) + d_H(\hat{\Theta}_I, \Theta_I) \xrightarrow{p} 0,
\]
thus (15) holds.

C. CDMCA is approximated by PMvGE with linear transformations

We argue an approximate relation between CDMCA and PMvGE, which is briefly explained in Section 3.6. In the below, we will derive the solution \( \hat{\psi}_{\text{CDMCA}} \) of a slightly modified version of CDMCA, and an approximate solution \( \hat{\psi}_{\text{AprPMvGE}} \) of PMvGE with linear transformations. We then show that these two solutions are equivalent up to a scaling in each axis of the shared space.

C.1. Solution of a modified CDMCA

The original CDMCA imposes the quadratic constraint (8) for maximizing the objective function (7). Here we replace \( w_{ij} \) in (8) with \( \delta_{ij} \) so that the constraint becomes
\[
\sum_{i=1}^{n} \psi(d_{ij})^\top x_i x_i^\top \psi(d_{ij}) = I.
\]
This modification changes the scaling in the solution, but the computation below is essentially the same as that in Shimodaira (2016). Let us define the augmented data vector, called “simple coding” (Shimodaira, 2016), \( \tilde{x}_i = (0_{p_1}, \ldots, 0_{p_{d_i-1}}, x_i, 0_{p_{d_i+1}}, \ldots, 0_{p_D}) \in \mathbb{R}^p \) where \( p := p_1 + p_2 + \cdots + p_D \). Now, data matrix is \( X := (\tilde{x}_1^\top, \tilde{x}_2^\top, \ldots, \tilde{x}_n^\top)^\top \in \mathbb{R}^{n \times p} \), and the parameter matrix is \( \psi := (\psi^{(1)}^\top, \psi^{(2)}^\top, \ldots, \psi^{(D)}^\top)^\top \in \mathbb{R}^{p \times K} \). With this augmented representation, the \( D \)-view embedding is now interpreted as a 1-view embedding. CDMCA maximizes the objective function
\[
\frac{1}{2} \sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij} \langle \psi(d_{ij})^\top x_i, \psi(d_{ij})^\top x_j \rangle = \operatorname{tr} (\psi^\top H \psi) \quad (H := X^\top W X),
\]
with respect to \( \psi \) under constraint \( \psi^\top G \psi = I \) where \( G := X^\top X \). Let \( U_K \) be the matrix composed of the top-\( K \) eigenvectors of \( G^{-1/2} H G^{-1/2} \). Then the solution of the modified CDMCA is
\[
\hat{\psi}_{\text{CDMCA}} := G^{-1/2} U_K.
\]
C.2. Approximate solution of PMvGE with linear transformations

MLE of PMvGE maximizes $\ell_n(\alpha, \psi)$ defined in (6). Here we modify it by adding an extra term as $\tilde{\ell}_n(\alpha, \psi) := \ell_n(\alpha, \psi) - \frac{1}{2} \sum_{i=1}^{n} \sum_{(d_i, d_j) \in D} h_{ij}(\alpha, \psi)$. The difference approaches zero for large $n$, because $|\ell_n(\alpha, \psi) - \tilde{\ell}_n(\alpha, \psi)|/|\ell_n(\alpha, \psi)| = O(n^{-1})$. Since the parameter $\alpha$ is not considered in CDMCA, we assume $D := \{\text{all pairs of views}\}$ and $\alpha = (\alpha^{(d,e)}), \alpha^{(d,e)} \equiv \alpha_0 > 0$ ($\forall d, e$). We further assume that the transformation of PMvGE is linear: $f_\psi(dx) = \psi^{(d)}(x)$ ($\forall d$), and data vectors in each view are centered. With this setting, we will show that the maximizer of a quadratic approximation of $\tilde{\ell}_n$ is equivalent to CDMCA.

To rewrite the likelihood function as $\tilde{\ell}_n(\alpha, \psi) = \frac{1}{2} \sum_{i=1}^{n} \sum_{j=1}^{n} S_{ij}(g_{ij}(\psi))$, we define $g_{ij}(\psi) := (\psi^{(d)}(x_i, \psi^{(d)}(x_j))$ and $S_{ij}(g) = w_{ij} \log(\alpha_0 \exp(g) - \alpha_0 \exp(g), g \in \mathbb{R}$. Since $S_{ij}(g)$ is approximated quadratically around $g = 0$ by

$$S_{ij}^Q(g) = \alpha_0 \left\{ -\frac{1}{2} g^2 + \left( \frac{w_{ij}}{\alpha_0} - 1 \right) g \right\} + S_{ij}(0),$$

$\tilde{\ell}_n(\alpha, \psi)$ is approximated quadratically by

$$\tilde{\ell}_n^Q(\psi) := \frac{1}{2} \sum_{i=1}^{n} \sum_{j=1}^{n} S_{ij}^Q(g_{ij}(\psi)) = \alpha_0 \left\{ -\frac{1}{2} \text{tr} \left( \left( \psi^T \Gamma \psi \right)^2 \right) + \frac{1}{\alpha_0} \text{tr} \left( \psi^T \Gamma \psi \right) - \text{tr} \left( \psi^T \Gamma \psi \right) \right\} + \frac{1}{2} \sum_{i=1}^{n} \sum_{j=1}^{n} S_{ij}(0), \quad (17)$$

where $G = X^T X$, $H = X^T W X$. The function $\tilde{\ell}_n^Q(\psi)$ has rotational degrees of freedom: $\tilde{\ell}_n^Q(\psi) = \tilde{\ell}_n^Q(\psi O)$ for any orthogonal matrix $O \in \mathbb{R}^{K \times K}$. Thus, we impose an additional constraint $\psi^T \Gamma \psi = \Gamma_K := \text{diag}(\gamma_1, \gamma_2, \ldots, \gamma_K)$ for any $(\gamma_1, \ldots, \gamma_K) \in \mathbb{R}^{K}_{\geq 0}$. $\psi$ satisfying this constraint is written as $\psi = G^{-1/2} V_K \Gamma_K^{1/2}$, where $V_K \in \mathbb{R}^{P \times K}$ is a column-orthogonal matrix such that $V_K^T V_K = I$. By substituting $\psi = G^{-1/2} V_K \Gamma_K^{1/2}$ into eq. (17), we have

$$\tilde{\ell}_n^Q(\psi) = \alpha_0 \left\{ -\frac{1}{2} \text{tr} \left( \Gamma_K^2 \right) + \frac{1}{\alpha_0} \text{tr} \left( \Gamma_K S_K \right) \right\} + \text{Const.} = \frac{\alpha_0}{2} \left\{ \frac{1}{\alpha_0} \| S_K \|_F^2 - \left\| \Gamma_K - \frac{1}{\alpha_0} S_K \|_F^2 \right\} \right\} + \text{Const.,} \quad (18)$$

where $S_K = V_K^T G^{-1/2} H G^{-1/2} V_K$ and $\| \cdot \|_F$ denotes the Frobenius norm. This objective function is maximized when $\Gamma_K = \frac{1}{\alpha_0} S_K$ and $V_K = U_K$, because $\min_{\Gamma_K} \| \Gamma_K - \frac{1}{\alpha_0} S_K \|_F^2 = 0$ is achieved by $\Gamma_K = \frac{1}{\alpha_0} S_K$, and $\max_{\Gamma_K} \| S_K \|_F^2$ is achieved by $V_K = U_K$ where $U_K$ is the matrix composed of the top-$K$ eigenvectors of $G^{-1/2} H G^{-1/2}$. Therefore, $\tilde{\ell}_n^Q(\psi)$ is maximized by

$$\hat{\psi}_{\text{Apr.PMvGE}} = G^{-1/2} U_K \Gamma_K^{1/2}.$$

By substituting $V_K = U_K$ into $\Gamma_K = \frac{1}{\alpha_0} S_K$, we verify that $\Gamma_K$ is a diagonal matrix with $\gamma_k := \lambda_k / \alpha_0$ where $\lambda_k$ is the $k$-th largest eigenvalue of $G^{-1/2} H G^{-1/2}$ ($k = 1, 2, \ldots, K$).

C.3. Equivalence of the two solutions up to a scaling

By comparing the two solutions, we have

$$\hat{\psi}_{\text{Apr.PMvGE}} = \hat{\psi}_{\text{CDMCA}} \Gamma_K^{1/2}. \quad (19)$$

This simply means that each axis in the shared space is scaled by the factor $\sqrt{\gamma_k}, k = 1, \ldots, K$. Let $\bar{y}, \bar{y}'$ be feature vectors in the shared space computed by the approximate PMvGE with linear transformations, and $y, y'$ be feature vectors in the shared space computed by the modified CDMCA. Then the inner product is weighted in PMvGE as

$$\langle \bar{y}, \bar{y}' \rangle = \sum_{k=1}^{K} \gamma_k y_k y'_k = \sum_{k=1}^{K} y_k y'_k.$$
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