Association between continuity of care in general practice and hospital admissions for ambulatory care sensitive conditions: cross sectional study of routinely collected, person level data

Isaac Barker, Adam Steventon, Sarah R Deeny

ABSTRACT

OBJECTIVE
To assess whether continuity of care with a general practitioner is associated with hospital admissions for ambulatory care sensitive conditions for older patients.

DESIGN
Cross sectional study.

SETTING
Linked primary and secondary care records from 200 general practices participating in the Clinical Practice Research Datalink in England.

PARTICIPANTS
230 472 patients aged between 62 and 82 years and who experienced at least two contacts with a general practitioner between April 2011 and March 2013.

MAIN OUTCOME MEASURE
Number of hospital admissions for ambulatory care sensitive conditions (those considered manageable in primary care) per patient between April 2011 and March 2013.

RESULTS
We assessed continuity of care using the usual provider of care index, which we defined as the proportion of contacts occurring between April 2011 and March 2013 that were with the most frequently seen general practitioner. On average, the usual provider of care index score was 0.61. Continuity of care was lower among practices with more doctors (average score 0.59 in large practices versus 0.70 in small practices). Higher continuity of care was associated with fewer admissions for ambulatory care sensitive conditions. When modelled, controlling for demographic and clinical patient characteristics, an increase in the usual provider of care index of 0.2 for all patients would reduce these admissions by 6.22% (95% confidence interval 4.87% to 7.55%). There was greater evidence for an association among patients who were heavy users of primary care. Heavy users also experienced more admissions for ambulatory care sensitive conditions than other patients (0.36 admissions per patient for those with ≥18 contacts with a general practitioner, compared with 0.04 admissions per patient for those with 2-4 contacts).

CONCLUSIONS
Strategies that improve the continuity of care in general practice may reduce secondary care costs, particularly for the heaviest users of healthcare. Promoting continuity might also improve the experience of patients and those working in general practice.

Introduction
Healthcare systems around the world are facing increases in unplanned hospital admissions.1 These admissions are often undesirable for patients, disruptive of elective care, and costly, so their prevention has become a priority. A large proportion of hospital admissions could be avoided through effective management and treatment in the community,2 which has become a focus of initiatives such as the accountable care organisations in the United States and the “five year forward view” strategy in the National Health Service in England.3 4 The focus has often been on improving access to primary care (for example, by increasing opening hours or introducing remote care).5 6 These initiatives have achieved mixed impacts on patient outcomes,7 and some may have had unintended effects on the continuity of the care provided.8-10

Although timely access is an important goal for high quality healthcare,11 continuity of care is associated with patient satisfaction,1213 as well as quality of life for patients living with long term conditions.14 It encompasses several aspects, including the consistency of care with a healthcare professional over time, quality of the interpersonal relationships between healthcare professionals and patients, and availability of information about the patient.12 General practitioners value continuity of care, often considering it to be a core value of their profession,15 16 while many patients value a personal doctor to coordinate and integrate their care.12 17 Unfortunately, there are indications from patient surveys that continuity of care is in decline in both the United States and England,18 19 at a time when doctors are reporting concerns about their ability to provide coordinated care for the increasing number of

WHAT IS ALREADY KNOWN ON THIS TOPIC
To decrease unplanned hospital admissions, many healthcare systems have focused on improving speed and equity of access to primary care
Continuity of care has been falling in England, perhaps because there are trade-offs between promoting access and continuity
Continuity of care is associated with patient and practitioner satisfaction, but its link with hospital admissions has been unclear

WHAT THIS STUDY ADDS
Continuity of care is associated with lower hospital admissions for ambulatory care sensitive conditions, which are those considered manageable in primary care
For heavy users of general practice, the association between continuity of care and hospital admissions for ambulatory care sensitive conditions is clearest
Strategies that improve continuity of care might simultaneously improve quality and reduce cost, although interventions will require evaluation
our study is that, as with some previous research, we have used person level data from the electronic medical record in primary care to measure continuity of care. These data are increasingly available for use in quality improvement and can be assembled in a timely fashion at minimal cost. Thus measurement strategies based on these data have widespread applicability. If these indices show a link with patient outcomes that can be improved through intervention in primary care, then they could support quality improvement activities.

Methods

Dataset

This was a retrospective, observational, and cross sectional study using data from the Clinical Practice Research Datalink, which contains the electronic medical records of participating general practices, and includes contacts with healthcare professionals, details of diagnostic tests, referrals to specialist healthcare, diagnosing health conditions, personal characteristics, and the start and end dates of patient registrations with general practices. Patients in the practices participating in the Clinical Practice Research Datalink are broadly representative of patients registered to practices in England.

The dataset was pseudonymised (all patient identifiable data had been removed) but linked at the person level to the hospital episode statistics, a national database containing administrative records from all hospitals in England that provide NHS care. From that database we obtained dates of inpatient admission and details of the diagnoses recorded during the inpatient stay (based on international classification of diseases, 10th revision, or ICD-10, codes). We also obtained and linked data on deaths occurring inside or outside of hospital from the Office for National Statistics, and socioeconomic deprivation scores for the small area of individual patient residence (index of multiple deprivation 2010).

Inclusion and exclusion criteria

We studied patients registered at general practices that participated in the Clinical Practice Research Datalink and that supplied data that met a minimum standard of quality. We included patients who were aged between 62 and 82 on 6 October 2013 and experienced at least two contacts with a general practitioner between April 2011 and March 2013—the period analysed in the study. To allow for a minimum of two years’ follow-up for every patient, we excluded those who died before March 2013. We also required that patients were registered with a Clinical Practice Research Datalink practice for at least one year before April 2011, so that we could derive the study covariates. Finally, we excluded a small number of patients whose age, index of multiple deprivation score, or sex was not recorded.

Admissions for ambulatory care sensitive conditions

Our outcome was the number of inpatient admissions for ambulatory care sensitive conditions for each patient between April 2011 and March 2013. Although several similar definitions of ambulatory care sensitive
admissions exist, after a review of the literature we chose a core set of 22 such conditions, which we adapted from Bardsley et al (see supplementary appendix 1). Using the recorded ICD-10 codes we identified the corresponding admissions from the hospital episode statistics data. For some conditions (eg, asthma) we limited our search to the primary diagnosis field, whereas for others (eg, gangrene) we included all fields (see supplementary appendix 1).

Continuity of care
Continuity of care is a complex, multifaceted concept reflecting at least four domains: the subjective experience of a caring relationship (interpersonal continuity), a history of interacting with the same healthcare professional across a series of discrete episodes (longitudinal continuity), the availability of clinical and psychosocial information across encounters and professionals (informational continuity), and the effective collaboration of teams across care boundaries to provide seamless care (management continuity). In this study, we dealt with longitudinal continuity of care, which we measured using the usual provider of care index. Usual provider of care is defined as the proportion of a patient’s contacts that was with their most regularly seen doctor. For example, if a patient had 10 general practitioner contacts, including six with the same doctor, then their usual provider of care index score would be 0.6. We measured the usual provider of care index score for April 2011 to March 2013, the same period used for hospital admissions.

Study covariates
For each patient we determined age, sex; socioeconomic deprivation score grouped into fifths (based on national thresholds, with the first fifth representing the least deprived), number of contacts with a general practitioner (April 2011 to March 2013), number of active previous long term health conditions (recorded before April 2011), and number of previous referrals to specialist care (in 2010/11).

We identified contacts with general practitioners from the Clinical Practice Research Datalink data by selecting records corresponding to specific types of encounter (eg, clinic, home visit, telephone) and which were entered by general practitioners (including partners, registrars, sole practitioners, and locum doctors). All contacts were linked to a specific practitioner. Contacts with a general practitioner were limited to one each day, owing to concerns about duplication. We also limited the number of referrals to specialist care in the preceding year to one each day, to avoid duplication arising from counting two Read codes being entered for the same referral.

We identified long term conditions that were recorded in the general practice data between 2001 and 2011 and had not been recorded as resolved. We did not rely on the hospital episode statistics data to assess the long term conditions of individual patients, since in such data, conditions are more likely to be recorded for patients who are admitted frequently. The long term conditions included are the 17 measured as part of the Quality Outcomes Framework for general practice in England (see table 1).

Statistical analysis
We used multivariable generalised linear regression to test the association between the number of admissions for ambulatory care sensitive conditions and the usual provider of care index score, controlling for the covariates. We selected the distribution for our regression model based on an assessment of Pearson residuals. Since over-dispersion was present (and also the variance of the number of admissions for ambulatory care sensitive conditions was greater than the mean), we preferred a negative binomial model (with a log-link) to a Poisson model. We selected covariates using the Akaike Information Criterion; this supported including all of our chosen covariates. We expected that two patients in the same practice were likely to have more similar rates of admission than two patients in randomly different practices. Therefore, our model included practice level random effects, but we carried out likelihood ratio tests to determine whether the practice level variation was statistically significant from zero.

We interpreted our model by exponentiating the coefficients to estimate incidence rate ratios for an incremental change in the usual provider of care index score. These incidence rate ratios were then scaled to correspond to an increase in the usual provider of care index score of 0.2—around 1 standard deviation—and interpreted as the percentage change in hospital admissions for ambulatory care sensitive conditions. Thus, the incidence rate ratio shows our estimate of the relative change in the number of admissions to result from an increase in the usual provider of care index score of 0.2, which might occur from any baseline level (for example, from 0.3 to 0.5 or from 0.6 to 0.8), and when holding the covariates constant. Since this model assumed a linear relation between the usual provider of care index score and admissions, we also fitted a more flexible model that classified patients as receiving low, medium, and high levels of continuity, rather than treating the usual provider of care index score as a continuous variable. We defined low continuity a priori as a usual provider of care index score between 0 and 0.4, medium between 0.4 and 0.7, and high between 0.7 and 1.

Subgroup analysis
We conducted subgroup analysis for higher and lower users of general practice care, in each case examining the association between the number of admissions for ambulatory care sensitive conditions and the usual provider of care index score. This was based on dividing the study population into five subgroups of approximately the same size, based on the number of contacts with a general practitioner that each patient experienced over two years. Within each of these subgroups, we estimated the association between the usual provider of care index score and the number of admissions for ambulatory care sensitive conditions using equivalent models to those we have described. In a similar manner, we examined whether the relationship between continuity of care and admissions differed by age group.
Sensitivity analysis
Several metrics have been used to measure continuity of care at a patient level. To ensure our findings were not sensitive to our choice of metric, we additionally performed analysis using the Bice-Boxerman continuity of care index. Compared with the usual provider of care index, which focuses on the concentration of contacts with a particular general practitioner, the Bice-Boxerman continuity of care index takes into account the dispersion and distribution of contacts a patient has over different general practitioners. We re-ran the model predicting admissions to hospital for ambulatory care sensitive conditions including the Bice-Boxerman continuity of care index as an exposure instead of the usual provider of care index. Since the Bice-Boxerman continuity of care index score requires a minimum of three contacts, our analysis was restricted to that subgroup of patients. The Bice-Boxerman continuity of care index score as, with the usual provider of care index score, is bounded between 0 and 1. We also calculated a Pearson correlation coefficient for the two indices, for those patients with three or more contacts.

Patient involvement
Patients were not involved in setting the research question or the outcome measures, nor were they involved in developing plans for design or implementation of the study. We received valuable feedback on the interpretation and presentation of our results from a patient perspective as part of the review process of The BMJ. As our data are effectively anonymised, we cannot disseminate the results of the research to participants directly, although the Clinical Practice Research Datalink keeps a bibliography of studies published using their data. We plan to disseminate the results to relevant patient communities through alternative mediums, such as blogs and press releases, and by working with national organisations.

Results
Overall, 263829 patients from 200 general practices met our inclusion and exclusion criteria, of whom 230472 had at least two contacts with general practitioners between April 2011 and March 2013 (again, from 200 practices). On average over the two years, each of these patients experienced 11.40 (SD 9.40) contacts with a general practitioner (table 1). The mean usual provider of care index score was 0.61 (SD 0.23); figure 1 shows the distribution.

The low continuity of care group comprised 52550 patients (22.8%), the medium group 96902 patients (42.1%), and the high group 81020 patients (35.2%). These groups were similar in terms of their age, sex, and socioeconomic deprivation score (table 1). However, patients in the low continuity group had more contacts with general practitioners on average than the other groups (13.11 per person, compared with 11.34 per person in the medium group and 10.37 in the high group).

Continuity of care as measured by the usual provider of care index tended to be lower in larger practices. In practices with between one and three full-time equivalent general practitioners, 49.7% of patients (n=11883) experienced high continuity of care according to this metric, compared with 30.7% of patients (n=34023) in practices with seven or more full-time equivalent

---

**Table 1 | Descriptive statistics for patients who experienced at least two contacts with general practitioners and for subsets of patients with low, medium, and high continuity of care. Values are numbers (percentages) unless stated otherwise**

| Variables                                                       | Continuity of care* | Medium (n=96902) | High (n=81020) | All patients (n=230472) |
|-----------------------------------------------------------------|---------------------|------------------|----------------|-------------------------|
| Women (% of patients)                                           | 55.77               | 53.95            | 51.88          | 53.63                   |
| Mean (SD) age (years)                                           | 71.15 (5.88)        | 71.38 (5.87)     | 71.66 (5.87)   | 71.43 (5.88)            |
| Socioeconomic deprivation fifth‡:                               |                     |                  |                |                         |
| First (least deprived)                                         | 15,479 (29.66)      | 26,274 (27.11)   | 18,906 (23.33) | 60,659 (26.32)          |
| Second                                                          | 12,360 (23.52)      | 24,796 (25.59)   | 21,359 (26.36) | 58,515 (23.39)          |
| Third                                                           | 9,903 (18.84)       | 19,607 (20.23)   | 17,958 (22.26) | 47,464 (20.59)          |
| Fourth                                                          | 8,253 (15.71)       | 15,733 (16.24)   | 14,403 (17.78) | 38,389 (16.66)          |
| Fifth (most deprived)                                          | 6,555 (12.47)       | 10,486 (10.83)   | 8,394 (10.36)  | 25,445 (11.04)          |
| Mean (SD) No of long term conditions per patient†              | 1.16 (1.18)         | 1.09 (1.14)      | 1.06 (1.14)    | 1.10 (1.14)             |
| No of long term conditions per patient‡:                       |                     |                  |                |                         |
| 0                                                               | 18,345 (34.91)      | 35,895 (37.04)   | 30,103 (37.16) | 84,343 (36.60)          |
| 1                                                               | 17,439 (33.19)      | 31,915 (32.94)   | 27,279 (34.22) | 77,083 (33.45)          |
| ≥2                                                              | 16,766 (31.90)      | 29,092 (30.02)   | 23,188 (28.62) | 69,046 (29.96)          |
| Mean (SD) No of contacts with general practitioner over two years per patient | 13.11 (10.06) | 11.34 (9.39) | 10.37 (8.80) | 11.40 (9.40) |
| Mean (SD) No of referrals to specialist care in 2010/11 per patient | 0.5 (0.94) | 0.46 (0.90) | 0.41 (0.83) | 0.45 (0.89) |
| Mean (SD) No of admissions for ambulatory care sensitive conditions per patient | 0.18 (1.09) | 0.15 (0.95) | 0.14 (0.97) | 0.16 (1.01) |
| No of admissions for ambulatory care sensitive conditions per patient: |                     |                  |                |                         |
| 0                                                               | 46,314 (88.09)      | 86,798 (89.53)   | 73,409 (90.57) | 206,521 (89.57)         |
| 1                                                               | 45,599 (86.67)      | 79,691 (83.73)   | 57,055 (70.4)  | 177,575 (77.0)          |
| ≥2                                                              | 17,023 (32.74)      | 26,562 (27.46)   | 19,934 (23.99) | 62,909 (27.31)          |
| Mean (SD) usual provider of care index score                    | 0.32 (0.06)         | 0.55 (0.08)      | 0.88 (0.11)    | 0.61 (0.23)             |

*Usual provider of care index scores: low continuity of care 0-0.4; medium continuity of care 0.4-0.7; high continuity of care 0.7-1.0.

‡Index of multiple deprivation; fifths based on national thresholds.

Included asthma, atrial fibrillation, cancer, coronary heart disease, chronic kidney disease, chronic obstructive pulmonary disease, dementia, depression, diabetes mellitus, epilepsy, heart failure, hypertension, rheumatoid arthritis, learning disabilities, mental health, stroke or transient ischaemic attacks, and peripheral arterial disease.
With medium continuity of care compared with people with low continuity of care, people associated with having more such admissions.

For patients with low continuity of care tended to have lower continuity of care, while also experiencing more admissions for ambulatory care sensitive conditions. Thus, the heaviest users of general practice had a mean usual provider of care index score of 0.56 (compared with 0.69 for the lowest fifth), and 0.36 admissions on average (compared with 0.04 for the lowest fifth) – see supplementary appendix 2. Table 5 presents the results of the subgroup analysis that examined how the association between the usual provider of care index score and admissions for ambulatory care sensitive conditions varied between fifths defined by general practice utilisation. There was some indication of a U-shaped curve, but the evidence for a link between the usual provider of care index score and admissions was greatest among the most frequent users of general practice. For patients in the two highest fifths of general practice utilisation (12-17 contacts and ≥18 contacts), a 0.2 increase in the usual provider of care index score was associated with a reduction in ambulatory care sensitive admissions of 3.32% (95% confidence interval 0.76% to 5.82%) and 3.97% (1.91% to 6.00%), respectively (P<0.001) (table 5). The associations were less statistically significant at lower levels of general practice utilisation. Table 5 also presents the results of the subgroup analysis for age group: although an inverse relation was observed between the usual provider of care index score and admissions for each age group, the strength of the relationship was larger among older patients.

Sensitivity analysis
We calculated the Bice-Boxerman continuity of care index score for the 214 777 patients who had at least three contacts with general practice in the study period. Among those patients, the Bice-Boxerman continuity of care...
Table 3 | Predictors of admissions for ambulatory care sensitive conditions to hospital among patients with a minimum of two contacts with a general practitioner (n=230 472)

| Predictors                          | Coefficient (95% CI) | P value |
|-------------------------------------|----------------------|---------|
| Usual provider of care index score  | −0.32 (−0.39 to −0.25) | <0.001  |
| Age (years)                         | 0.02 (0.02 to 0.02)   | <0.001  |
| No of pre-existing long term conditions per patient | 0.22 (0.21 to 0.23) | <0.001  |
| No of contacts with general practitioner over two years | 0.06 (0.06 to 0.06) | <0.001  |
| Sex (female v male)                 | −0.25 (−0.27 to −0.22) | <0.001  |
| Socioeconomic deprivation score by fifths*: |                      |         |
| Second versus first                 | 0.02 (−0.02 to 0.07)  | 0.26    |
| Third versus first                  | 0.05 (0.00 to 0.10)   | 0.04    |
| Fourth versus first                 | 0.07 (0.01 to 0.12)   | 0.01    |
| Fifth versus first                  | 0.08 (0.02 to 0.14)   | 0.01    |
| No of referrals to specialist care 2010/11 | 0.08 (0.06 to 0.09) | <0.001  |

*From first fifth (least deprived) to fifth fifth (most deprived).

In this model, the usual provider of care index score was entered as a continuous variable, with the association with admissions for ambulatory care sensitive conditions estimated using the negative binomial distribution with a log-link. Results can be interpreted by constructing incidence rate ratios for an incremental change in usual provider of care index score. As the variables have a log-link, those incidence rate ratios can be constructed for a unit change in a covariate by taking the exponential of the coefficients.

Discussion

Preventing hospital admissions is a priority in many countries, and in England unplanned admissions for ambulatory care sensitive conditions (those considered to be manageable in primary care) accounted for £1.42bn of spending in 2009/10 (£170 590 for each general practice). Since strategies are needed to manage demand for secondary care, we examined the association between continuity of care and hospital admissions among older patients (aged 62-82 years). Continuity of care varied considerably across general practices in England, with 48.6% of patients in smaller practices and 30.7% of patients in larger practices falling into our defined high continuity group.

Patients who saw the same general practitioner a greater proportion of the time experienced fewer admissions to hospital for ambulatory care sensitive conditions than other patients. Indeed, compared with patients with low continuity of care, patients with medium continuity of care experienced 8.96% fewer of these admissions, and those with high continuity of care experienced 12.49% fewer. There was greater evidence for an association between continuity of care and admissions among those with the highest levels of contacts with general practitioners. As well as showing this clearer link between continuity and reductions, the heaviest users had lower continuity of care (mean usual provider of care index 0.56 v 0.61) and more admissions for ambulatory care sensitive conditions (0.36 per person on average v 0.16).

Our findings suggest there is opportunity to improve care and reduce hospital admissions among the highest users of healthcare, but there are various explanations of our findings that must first be carefully considered. One possibility is that when patients receive more continuous care, their doctors might be better able to understand their health needs within time constrained appointments and thus provide care that is more appropriate to the needs of the patient. Continuity of care score was strongly correlated with the usual provider of care index score, showing a Pearson correlation coefficient of 0.96. When we repeated our modelling using the Bice-Boxerman continuity of care index, a 0.2 increase in that score was associated with 3.58% fewer admissions for ambulatory care sensitive conditions (95% confidence interval 2.60% to 4.76%, P<0.001).

Table 4 | Relation between admissions for ambulatory care sensitive conditions and continuity of care (n=230 472)

| Change in continuity of care | Relative % change in admissions for ambulatory care sensitive conditions (95% CI) | P value |
|-----------------------------|----------------------------------------------------------------------------------|---------|
| Medium (versus low)         | −8.96 (−5.63 to −14.22)                                                          | <0.001  |
| High (versus low)           | −12.49 (−9.45 to −19.29)                                                         | <0.001  |
| High (versus medium)        | −3.87 (−4.30 to −7.19)                                                           | 0.03    |

In this model, the usual provider of care index score was entered as an ordered, categorical variable (with low, medium, and high values), with admissions for ambulatory care sensitive conditions estimated using the negative binomial distribution with a log-link. The model was used to control for age, number of pre-existing long term conditions, number of contacts with a general practitioner, sex, socioeconomic deprivation, and number of referrals to specialist care in preceding year.

Table 5 | Estimated effect of a 0.2 increase in usual provider of care index score on admissions for ambulatory care sensitive conditions, for subgroups of patients

| Group or subgroup of analysis | Relative % change in admissions for ambulatory care sensitive conditions (95% CI) for a 0.2 increase in usual provider of care index score | P value |
|------------------------------|--------------------------------------------------------------------------------------------------------------------------|---------|
| All patients (n=230 472)     | −6.22 (−7.55 to −6.87)                                                                                                | <0.001  |
| Fifths of contact (No of contacts) |                                                                       |         |
| First (2-4) (n=49161)        | −3.48 (−7.65 to 0.87)                                                                                                | 0.12    |
| Second (5-7) (n=47 208)      | −3.89 (−7.49 to −0.14)                                                                                                | 0.04    |
| Third (8-11) (n=48 617)      | 0.34 (−2.59 to 3.36)                                                                                               | 0.82    |
| Fourth (12-17) (n=43 246)   | −3.32 (−5.82 to −0.76)                                                                                                | 0.01    |
| Fifth (21+) (n=42 240)       | −3.97 (−6.00 to −1.91)                                                                                                | <0.001  |
| Age group (years) of patients in 2012 with at least two contacts with general practice: |                                       |         |
| 62-66 (n=10 761)             | −3.40 (−5.57 to −1.19)                                                                                                | 0.003   |
| 69-75 (n=7 461)              | −7.92 (−10.22 to −5.65)                                                                                                | <0.001  |
| 76-82 (n=5 550)              | −8.37 (−10.71 to −5.99)                                                                                                | <0.001  |

In these models, the usual provider of care index score was entered as a continuous variable, with admissions for ambulatory care sensitive conditions using the negative binomial distribution with a log-link. Results were interpreted by constructing incidence rate ratios for an incremental change in usual provider of care index score. As the variables have a log-link, incidence rate ratios for a unit change in a covariate were constructed by taking the exponential of the coefficients. The results were interpreted for a 0.2 change of the usual provider of care index score. Each model is adjusted to control for age, number of pre-existing long term conditions, number of contacts with a general practice, sex, socioeconomic deprivation, and number of referrals to specialist care in preceding year.
might also promote a more effective and trusting relationship between patients and doctors, leading to a better understanding of health problems and greater adherence to the agreed course of treatment. One would expect these benefits to be felt most keenly in those patients who need to visit the general practitioner frequently. However, it is important to note that identifying a causal mechanism is beyond the scope of this study.

To our knowledge no previous study has examined the association between longitudinal continuity of care in primary care (measured by usual provider of care index) and admissions for ambulatory care sensitive conditions. However, a study of older patients in primary care in one Canadian province found that those with a medium degree of longitudinal continuity of care, measured through the usual provider of care index, had 27% (95% confidence interval 25% to 29%) more visits to an emergency department than those with high continuity of care. And although it relates to a different process, the GP patient survey in England asks patients whether they were able to see the general practitioner of their choice, and practices scoring more highly on this metric tend to have lower rates of hospital admissions. A body of literature has looked at the effect of continuity in secondary care.

**Strengths and limitations of this study**

We were able to study linked records for 230,472 patients, resulting in small confidence intervals around our estimated values. Patients registered with practices participating in the Clinical Practice Research Datalink are broadly representative of those in England, but we restricted our analysis to practices submitting up-to-standard data. We examined the period between 2011 and 2013; however, we do not believe that the associations found are likely to have changed with time. We cannot directly assess the representativeness of the 200 practices in our study sample, although we note that 26.3% of our study sample lived in the 20% least socioeconomically deprived areas, compared with 11.0% in the 20% most deprived areas. This may be partly because our study was restricted to those aged between 62 and 82 years. We focused on older patients because they tend to experience more admissions for ambulatory care sensitive conditions than other age groups, but older people may be less likely to live in deprived areas. In addition, we limited our analysis to those patients with a minimum of two contacts with general practitioners, to ensure our metric was sensitive. Although we cannot establish the association between the usual provider of care index score and admissions for people younger than 62 or older than 82, we note that our subgroup analysis showed a strong age gradient within our sample. More studies are needed, but it is possible that continuity of care is more directly related to admissions among those with a higher burden of chronic disease, which includes older people as well as those living in more deprived areas.

We chose the usual provider of care index since it is more easily interpreted than some of the other metrics of care continuity. However, we performed sensitivity analysis using the Bice-Boxerman continuity of care index, which compared with the usual provider of care index is more sensitive to the distribution of contacts across multiple doctors (in addition to the proportion of contacts with most regular doctors). The two indices were highly correlated. Although the level of association differed, both showed an association with admissions for ambulatory care sensitive conditions. Furthermore, the association remained when we modelled the usual provider of care index score as a categorical rather than continuous variable. In that analysis, the thresholds we applied to divide patients into high, medium, and low continuity of care groups were necessarily arbitrary, and there were also some limitations to the way in which we calculated the usual provider of care index score. By focusing on general practitioners, we did not assess the importance of the continuity of care provided by practice nurses and healthcare assistants. Also, some caution is needed in relation to larger practices, since it is possible that in those practices patients are deliberately managed by several general practitioners working in a small team (for example, to provide cover when doctors work part time). Ideally, we would have adapted the usual provider of care index to allow for these “micro-teams,” but the requisite information was not available.

We aimed to assess longitudinal continuity of care, and we could not assess the quality of the interactions from the patient perspective or whether patients saw their preferred doctor. It is possible that measures of interpersonal, management, and informational continuity would also show an association with utilisation of secondary care, and that patients’ definitions of continuity of care might vary from that presented here. By analysing admissions for ambulatory care sensitive conditions, rather than all unplanned admissions, we aimed to focus on those admissions most likely to be preventable by good quality primary care. However, administrative data provide limited information to determine whether individual admissions are avoidable or desirable. For example, heart failure is often classified as an ambulatory care sensitive condition, but even when managed according to evidence informed guidelines, it will gradually deteriorate to the point at which hospital admission may be required. However, if only a subset of the admissions for ambulatory care sensitive conditions were avoidable, then this would bias the magnitude of our results towards showing a smaller association. We did not examine the association between continuity of care and other forms of secondary care, community healthcare, or social care, and indeed there is some evidence that increased continuity of care may lead to later referral for cancer care.
mechanism or direction of action—for example, general practitioners may alter their management of a patient immediately before or after an admission to hospital, potentially leading to improved continuity of care after hospital discharge. Moreover, although we examined patient subgroups defined by general practice utilisation, continuity of care might be particularly important for groups that we could not isolate from the data, such as those unable to adopt self care strategies. Continuity of care might also be more important at some times than others—for example, according to the onset of acute health problems or after discharge from hospital. Finally, the importance of continuity of care might vary according to local context and the broader approach to managing patients within a general practice, including the role of practice nurses and specialist general practitioners, how information is shared within the practice, and working in a micro-team.

Conclusions and implications
This study motivates a renewed focus on promoting continuity of care, and it suggests that continuity is an important consideration when designing approaches to reduce hospital admissions. Over the past decade, initiatives to prevent hospital admissions have tended to prioritise speed and equity of access to primary care over continuity of care. In England, one of the few recent national initiatives to improve the continuity of care in general practice was the mandatory introduction of named accountable general practitioners, but this had no discernible impact on longitudinal continuity of care, at least for patients aged 75 over the first nine months. Therefore, approaches are needed to improve continuity of care in general practice.

Several approaches have been suggested, including increasing patient awareness of the importance of continuity, changing receptionists’ behaviour or practice booking systems to promote continuity, and organising large practices into small teams, each of which care for a subset of the patients registered at the practice, meaning that patients receive continuous care at team level rather than individual level. These can be explored further. Of course it cannot be assumed on the basis of this study that interventions that successfully improve the continuity of care will necessarily lead to reductions in hospital admissions for ambulatory care sensitive conditions, and interventions will require careful evaluation. A complementary approach would be to make data on continuity of care available to general practice teams for quality improvement initiatives, as recommended by the Royal College of General Practitioners. In this study we measured continuity of care using data that are already available to general practice teams from the electronic medical record, although in some cases the analytical capability to analyse these data may be lacking.

Despite the limitations of these data, we have demonstrated a link with hospital admission rates. We have also shown that interventions could be targeted at easily identifiable groups of patients. This suggests that the usual provider of care index could inform strategies for quality improvement, although it may need to be adapted for local context, and its responsiveness to service change would need to be tested.

Continuity of care aligns closely with the reasons many people chose to work in general practice, is important to patients, correlates with outcomes, and shows variability between general practices. As we show, continuity of care is also associated with avoidable hospital admissions. Thus initiatives that improve the continuity of care have the potential to improve the quality of healthcare while reducing cost, both fundamental aims of healthcare systems.

We thank Therese Lloyd for helping to prepare the dataset for analysis, sharing code, and methodological advice; Will Warburton and Martin Roland for their helpful and insightful comments on an earlier version of this manuscript; and the four peer reviewers for feedback, which helped us to improve the manuscript. This study is in part funded by data from the Clinical Practice Research Datalink obtained under license from the UK Medicines and Healthcare products Regulatory Agency. However, the interpretation and conclusions contained in this study are those of the authors alone.

Contributors: IB, AS, and SD had the idea for the study and submitted a major amendment to a previous Clinical Practice Research Datalink application to facilitate the research. IB, AS, and SD came up with the statistical analysis plan. IB and SD carried out the analysis. IB, SD, and AS drafted and finalised the paper. IB is the guarantor.

Funding: None received.

Competing interests: All authors have completed the ICMJE uniform disclosure form at www.icmje.org/coi_disclosure.pdf and declare: all authors had financial support from The Health Foundation for the submitted work; no financial relationships with any organisations that might have an interest in the submitted work in the previous three years; no other relationships or activities that could appear to have influenced the submitted work.

Ethical approval: This study was approved by the Clinical Practice Research Datalink independent scientific advisory committee. This was secondary analysis of data submitted to the Clinical Practice Research Datalink, no patient consent forms were required to access this dataset.

Data sharing: The data controller of the data analysed is the Clinical Practice Research Datalink. Patient level data are available subject to their information governance requirements.

Transparency: The lead author (IB) affirms that the manuscript is an honest, accurate, and transparent account of the study being reported, that no important aspects of the study have been omitted, and that any discrepancies from the study as planned (and, if relevant, registered) have been explained.

This is an Open Access article distributed in accordance with the Creative Commons Attribution Non Commercial (CC BY-NC 3.0) license, which permits others to distribute, remix, adapt, build upon this work non-commercially, and license their derivative works on different terms, provided the original work is properly cited and the use is non-commercial. See: http://creativecommons.org/licenses/by-nc/3.0/.

1. Huntley A, Lasserson D, Wye L, et al. Which features of primary care affect unscheduled secondary care use? A systematic review. BMJ Open 2014;4:e004746. doi:10.1136/bmjopen-2013-004746.
2. Roland M, Abel G. Reducing emergency admissions: are we on the right track? BMJ 2012;345:e6017. doi:10.1136/bmj.e6017.
3. NHSE England. Five Year Forward View. www.england.nhs.uk/wp-content/uploads/2014/10/fyfv-web.pdf.
4. Whittington JN, Nolan K, Lewis N, Torres T. Pursuing the Triple Aim: The First 7 Years. Milbank Q 2015;93:263-300. doi:10.1111/1468-0009.12122.
5. Gowing TE, Cecil E V, Sojak MA, et al. Access to primary care and visits to emergency departments in England: a cross-sectional, population-based study. PLoS One 2013;8:e6-11. doi:10.1371/journal.pone.0066699.
6. Osborn R, Moulds D, Schneider EC, Dity MM, Squires D, Sannik DO. Primary care physicians in ten countries report challenges caring for patients with complex health needs. Health Aff (Millwood) 2015;34:2104-12. doi:10.1377/hlthaff.2015.1018.
7. Edes T, Kiosonas B, Vackovic NH, Nichols LD, Becker MM, Hossain M. Better access, quality, and cost for clinically complex veterans with home-based primary care. J Am Geriatr Soc 2014;62:1954-61. http://onlineibrary.wiley.com/doi/10.1111/jgs.13030/full. doi:10.1111/jgs.13030.
8. The lead author (IB) affirms that the manuscript is an honest, accurate, and transparent account of the study being reported, that no important aspects of the study have been omitted, and that any discrepancies from the study as planned (and, if relevant, registered) have been explained.

This is an Open Access article distributed in accordance with the Creative Commons Attribution Non Commercial (CC BY-NC 3.0) license, which permits others to distribute, remix, adapt, build upon this work non-commercially, and license their derivative works on different terms, provided the original work is properly cited and the use is non-commercial. See: http://creativecommons.org/licenses/by-nc/3.0/.
