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Stature and long-term labor market outcomes: Evidence using Mendelian randomization

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Highlights

- Taller people have higher earnings.
- We use genetic instruments for height.
- The OLS models show that 10 cm of extra height increases earnings by 13%.
- The IV point estimate is lower at 9% and not statistically significant.

Abstract

We use the Young Finns Study (N=∼2000) on the measured height linked to register-based long-term labor market outcomes. The data contain six age cohorts (ages 3, 6, 9, 12, 15 and 18, in 1980) with the average age of 31.7, in 2001, and with the female share of 54.7. We find that taller people earn higher earnings according to the ordinary least squares (OLS) estimation. The OLS models show that 10 cm of extra height is associated with 13% higher earnings. We use Mendelian randomization, with the genetic score as an instrumental variable (IV) for height to account for potential confounders that are related to socioeconomic background, early life conditions and parental investments, which are otherwise very difficult to fully account for when using covariates in observational studies. The IV point estimate is much lower and not statistically significant, suggesting that the OLS estimation provides an upward biased estimate for the height premium. Our results show the potential value of using genetic information to gain new insights into the determinants of long-term labor market success.

Keywords: Height; Stature; Height premium; Earnings; Employment
Introduction

Taller people reap higher earnings. This empirical finding has been documented in several studies. There are three main explanations for the labor market premium in height (e.g., Sargent and Blanchflower, 1994; Judge et al., 2004; Persico et al., 2004; Case and Paxson, 2008; Tao, 2014; Sohn, 2015; Yamamura et al., 2015). First, height is associated with cognitive skills (Case and Paxson, 2008). Second, non-cognitive skills, such as social skills, may play a role in the height premium (Persico et al., 2004). Based on these two explanations, height is related to other individual qualities, such as cognitive or non-cognitive skills. Third, there may also be other social explanations for the height premium, such as discrimination against short people in the labor market (e.g., Cinnirella and Winter, 2009) as a form of social-perceptual bias by which tall individuals are perceived to have more positive qualities irrespective of their true qualities (Hamstra, 2014).

Most empirical studies treat height as an exogenous variable when examining the link between height and labor market outcomes. However, there may be important confounders, such as socioeconomic background, early life conditions and parental investments, that influence both a person’s height and subsequent labor market outcomes. It is challenging to adequately account for the combined effect of these factors when

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1 Hübler (2016) summarizes the relevant literature.

2 There are also empirical studies that point to the role of muscular strength as an explanation for the height premium (Lundborg et al., 2014). Böckerman et al. (2010) find only limited evidence for this view in the Finnish setting.

3 Case and Paxson (2006), and Vogl (2014) treat height as endogenously determined. Height is influenced by childhood nutrition, childhood environment and the prevalence of childhood diseases.
using the covariates that are available in observational studies. For example, parental investments are notoriously difficult to comprehensively and accurately measure. Thus, the causal effect of height on earnings and labor market success largely remains an open question.

The literature has pursued two approaches to address causal effects between height and earnings. First, some empirical studies have used a twin design. With twin data, it is possible to eliminate shared environmental factors, such as the family background, neighborhood and peer effects, and genetic factors (e.g., Böckerman and Vainiomäki, 2013). Second, two earlier studies used genetic instruments for height (von Hinke Kessler Scholder et al., 2013; Tyrrell et al., 2016). Genetic information could be helpful because genetic markers that are correlated with height should not directly affect the outcome variable of interest (i.e. earnings or employment). The specific instrument used in this paper is based on the findings in the genetics literature. There is substantial heritability for body height (Silventoinen et al., 2003). However, the contribution of individual genetic variants is modest. As a result, we used a genetic score with variants that genome-wide association studies (GWASs) have found to be significantly associated with height in extensive population samples (Allen et al., 2010), minimizing the weak instrument problem.

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4 Under the assumption of time-invariant parental investments they can be accounted for using fixed effects in panel data if the explanatory variable is time-varying. However, because adult height is time-invariant, its effect would also be eliminated by using fixed effects. Nevertheless, twin data can be used as discussed in the next paragraph.

5 The usage of genetic instruments is known as “Mendelian randomization” in the medical literature. The basic idea of Mendelian randomization is that genetic factors are distributed randomly in the population so that genetic risk factors are independent of potential confounding factors. We explain this idea in detail in the next section of the paper.
We use administrative information on long-term labor market outcomes (earnings and labor market attachment). We chose this approach because cross-sectional measures of labor market outcomes are inaccurate proxies for individuals’ lifetime labor market attachment and earnings (Böhlmark and Lindquist, 2006). Moreover, the use of the comprehensive register-based, long-term measures reduces measurement error from non-response and reporting biases.

Our contribution to the sparse empirical literature on the effects of height using genetic information builds on the fact that von Hinke Kessler Scholder et al. (2013) did not examine the labor market outcomes and that Tyrrell et al. (2016) used a self-reported categorical annual household income from a single year. In contrast, our paper uses linked data with administrative information and focuses on earnings that are a better measure of labor market success than annual household income that is confounded by social income transfers and spouse’s income.

Methods

*Mendelian randomization*

Mendelian randomization refers to empirical studies that use genetic instruments to estimate the causal effects of exposure variables or traits in non-experimental (observational) data because it is often difficult or impossible to use randomized controlled trials (Tyrrell et al., 2016). The need for instrumentation arises from the presence of confounding factors that correlate with both the exposure and the outcome variable. This leads to bias in OLS estimation.
Figure 1 depicts the various effects and estimators in this setting. The IV or Wald estimator avoids the bias if the following conditions are fulfilled: (1) the genetic instrument (G) must correlate with the exposure variable (X), i.e. it must be informative; (2) the genetic instrument (G) must affect the outcome (Y) only through its effect on the exposure (X), i.e. the instrument must be exogenous; and (3) the instrument (G) and confounder (Z) must be independent, i.e. the genotype should not be associated with the confounding relations between (X) and (Y). The first condition is justified based on genome-wide association studies that provide evidence for the correlations between genotypes and exposure variables that in our research design is the height of persons, which we discuss in greater detail below. The second condition is essentially an assumption, which is strictly untestable, but some indirect evidence for its validity is provided below. The validity of the third condition is based on Mendel’s second law (independent assortment), which states that genotypes are assigned randomly when passed across generations. This implies that in the population, or a representative sample of the population, genotypes are distributed independently of any confounding factors, which is what gives the method its name. It achieves the randomization of the exposure via “nature” rather than controlled trial.

Data sources and the sample

We use a longitudinal research design based on linked data. The data on height and genetic markers are from the Cardiovascular Young Finns Study (YFS). The YFS began in 1980, when 4,320 participants in six age cohorts (ages 3, 6, 9, 12, 15 and 18 years) were randomly chosen from five Finnish university regions using the national population

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6 The YFS is described at [http://youngfinnsstudy.utu.fi/studydesign.html](http://youngfinnsstudy.utu.fi/studydesign.html)
register (Raitakari et al., 2008). A total of 3,596 people participated in the study in 1980, and seven follow-up studies have been conducted; the most recent was in 2011/12.  

To obtain register information on labor market outcomes, we linked the YFS to the Finnish Longitudinal Employer-Employee Data (FLEED) of Statistics Finland (SF) using unique personal identifiers. The matching was exact, and there were no misreported ID codes. FLEED includes information on individuals’ earnings and labor market attachment, which is taken directly from comprehensive administrative registers that are maintained by SF.

To account for key observable differences in the parental background, we linked the YFS to the Longitudinal Population Census (LPC) of SF from the year 1980. We used indicators for the parents’ university-level education as family background variables.

**Measures**

We used both earnings and labor market attachment as labor market outcomes. Register-based, long-term earnings were measured as the average wage and salary earnings over the period 2001-2012. Labor market attachment was measured as the average employment years over the period 2001-2012.

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7 Written informed consent was obtained from participants who were at least 9 years old and from the parents of younger participants. The research plan and data collection procedures were accepted by the participating universities review boards, and data collection was conducted according to WHO standards as well as the Helsinki Declaration.
Height was measured in 2001, when the participants were 24 to 39 years old. Height was collected in professional health examinations that were conducted at local health centers. Thus, the measure contains a minimal amount of measurement error.

We used the genetic score as an instrument for height. The score was calculated from data for the 180 single nucleotide polymorphisms (SNPs) that are significantly (p<5 \times 10^{-8}) associated with height on the genome wide level according to Allen et al. (2010). A non-weighted genetic risk was calculated as the sum of genotyped risk alleles or imputed allele dosages carried by an individual. A Kernel density plot of the genetic score that is used in the IV estimations is presented in Appendix A2.

Genotyping was performed using a custom-made Illumina Human 670K BeadChip. Genotypes were called using the Illumina clustering algorithm (Teo et al., 2007). After quality control, there were 2442 samples and 546677 genotyped SNPs available for further analysis. Imputation was performed with MACH software and HapMap release 22 as a reference panel.

The genetic score has two major advantages. First, it is more powerful than any of the individual SNPs because it explains more variation in height. Second, it is more valid because it reduces the risk that any individual single nucleotide polymorphism will bias

More information on the SNPs and alleles used in calculating the risk score is presented in Appendix A1. We also used a weighted score in a robustness check below. A weighted genetic score was calculated as a sum of imputed allele dosages carried by an individual each multiplied by the effect size reported by Allen (2010). The weighted sum of effect alleles was divided by the mean effect size and transformed to the z-score. Effect alleles, effect sizes and the imputation quality of each variant used in the gene score calculation are reported in Appendix A1.
the IV estimates via an alternative biological pathway (pleiotropy) (Palmer et al., 2012). An alternative method to mitigate the bias caused by pleiotropy is to use individual SNPs with the Egger method (Bowden et al., 2015).

A genetic risk score for blood pressure was calculated using 29 SNPs associated with systolic and/or diastolic blood pressure in a genome-wide association study (International Consortium for Blood Pressure Genome-Wide Association Studies, 2011). In another genome-wide association study, genetic risk scores for total cholesterol and triglycerides were calculated with 25 SNPs associated with the total cholesterol concentration and 24 SNPs associated with the triglyceride concentration (Teslovich et al., 2010). All genetic scores were calculated as the sum of genotyped risk alleles or imputed allele dosages carried by an individual.

A computerized cognitive testing battery (CANTAB®) was used to assess cognitive performance in 2026 participants in the latest YFS follow-up study in 2011. The cognitive test battery included the following five tests: 1) the motor screening test (MOT), which was used as a training/screening tool to indicate difficulties in test execution; 2) the paired associates learning test (PAL), which measures visual and episodic memory and visuospatial associative learning; 3) the spatial working memory test (SWM), which measures short-term and spatial working memory as well as problem solving; 4) the reaction time test (RTI), which measures reaction, movement speed and attention; and 5) the rapid visual information test (RVP), which measures visual processing, recognition and sustained attention. Each test produced several variables.

We use the standard econometric methods of economics because the relevant literature (in economics) has used similar methods also. Otherwise it would be difficult to compare our estimation results to the earlier ones.

9
Principal component analysis was conducted to identify components accounting for the majority of the variation of the dataset (see Appendix A3). Principal component analyses were also performed separately for all individual tests. The first components resulting from these analyses were considered to represent cognitive performance related to the particular cognitive domain. The component for the motor screening test was excluded from further analysis, as it did not discriminate the subjects, which indicated a ceiling effect. Other components were normalized using the rank order normalization procedure, resulting in four variables, each with a mean of 0 and standard deviation (SD) of 1. All available data for each cognitive test were used in the analyses; therefore, the number of participants varied between the models (data on PAL and RTI tests were excluded for N=177 participants due to technical problems with the test equipment, while N=51 refused to participate in all or some of the tests). Detailed descriptions and validation of the cognitive data have been reported previously (Rovio et al., 2016).

Personality characteristics were assessed in 1983 for the four oldest cohorts using the Hunter-Wolf A-B Rating Scale (Wolf et al., 1982). The Hunter-Wolf Rating scale consists of four components (Aggression, Leadership, Responsibility and Eagerness-Energy) that were measured with 3-8 items. Responses to the items were given on a 7-point scale (1=‘totally disagree’ and 7=‘totally agree’) (Jokela and Keltikangas-Järvinen, 2009; Hintsa et al., 2014).

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10 We used SAS and the PROC FACTOR command in the analysis.

11 The Young Finns Study does not contain information on Big Five personality traits that started to gain popularity in the later part of the 1980s. The Hunter-Wolf rating scale was designed to measure behavior. Thus, instead of intrinsic tendencies, our personality characteristics measure an individual’s typical behavioral patterns that arguably have a significant role as a determinant of labor market success.
There are two main elements in our analyses. First, we run OLS models to replicate standard observational studies. Second, we estimate instrumental variable models wherein height, in 2001, is instrumented using the genetic score, providing the causal effects of height on the long-term labor market outcomes. We control for birth month, birth year effects and gender. These factors potentially impact earnings and employment, but they are predetermined with respect to our outcome variables and independent of the height risk factor.

We also include the parental education level to control for any intergenerational correlation of socioeconomic status (cf. Öberg, 2014). This accounts for possible omitted variable bias from the socioeconomic status of parents, which could affect the genetic risk factors, such as through assortative mating within educational groups. This would violate the independence assumption of the IV design, as argued in von Hinke et al. (2016). The problem can be corrected by conditioning on parental education, particularly the mother’s education level, which varies significantly between high and low risk factor individuals (Table 3).

Results

Descriptive evidence indicates that taller people have substantially higher earnings (Table 1). The baseline OLS estimates (Table 2, Panel A) show that height is statistically and economically significantly associated with average earnings over the period 2001-
2012. The quantitative size of the estimate is considerable. The coefficient for height implies that 10 cm of extra height is associated with 13% higher earnings. The estimate is roughly comparable to the earlier estimates that have been reported using Finnish data (Böckerman et al., 2010; Böckerman and Vainiomäki, 2013). The OLS estimate for labor market attachment is also statistically significant. The quantitative size of the effect corresponds to four months of additional employment during the twelve years covered by the data.

Next, we turn to the preferred IV estimates (Table 2, Panel B). The genetic score is a powerful instrument for height. In the first stage of IV, the F statistic on the instrument is 188 in the earnings regression, which exceeds the minimum standard of F=10 suggested in Staiger and Stock (1997) by a wide margin. Based on McClellan et al. (1994), we divide our estimating sample into those with an above-average value and those with a below-average value for genetic score and evaluate whether the two groups significantly differ in their observable characteristics that conceivably correlate with the second-stage outcome (Table 3). It is impossible to prove the null hypothesis that the instrument is uncorrelated with the second-stage error term. However, the lack of correlation between the instrument and observed variables, as shown in Table 3, is consistent with the exclusion restriction.  

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13 We also tested the exclusion restriction by dividing the sample into three ranges based on the gender-specific risk score. Analogously to the Goldfeld and Quandt (1965) test for homoscedasticity, we omitted a quarter of the central observations and used the individuals in the remaining lowest 3/8 and highest 3/8 to evaluate whether these groups significantly differed in their observable characteristics. We found significant differences (at least at the 10% level) in the following characteristics: height in 2001, visual and episodic memory, leadership, annual income (father), high education (mother) and blood pressure. Note that we control for these additional characteristics in Table 4.
The IV point estimate (Table 2, Panel B) for the effect of height on earnings is 9% per 10 cm of extra height, and the estimate is no longer statistically significant. The effect of height on average employment years over the period 2001-2012 is also much lower and insignificant in the IV estimation. The earlier Finnish evidence on the amount of the height premium is very limited. Böckerman and Vainiomäki (2013) report a 3.3% earnings effect per 1 cm for men and a 1.4% effect for women, but the effect remains significant only for women when using twin-differences and IV estimation. Using genetic instrument, Tyrrell et al. (2016, p. 6) report that “… a genetically determined 1 SD (6.3 cm) greater height was associated with a 0.05 (0.03 to 0.07) SD increase in annual household income.” Comparing this effect to our result is difficult because their annual household income is measured using five categories and because they do not report the SD for it.

An important finding of our research setting is that the IV estimates suggest that the OLS estimates are potentially upward biased. For example, specific early life conditions and parental investments that are not accounted for in the OLS models may lead to both greater height through differences in the children’s nutrition and health and thereby to better long-term labor market outcomes later in life. This unaccounted confounder may explain the higher point estimates for height in the OLS estimation.

A caveat to the finding of a null effect for height in the IV setting is that it may be a symptom of low power due to an insufficient sample size. Freeman et al. (2013) show that the sample size needed in Mendelian randomization study is inversely proportional to two factors: the variation in the exposure or observable trait explained by the genetic instrument, and the square of the true effect size. Genome-wide association studies usually report low shares of explained variation for genotypes; therefore, the minimum
sample size required for detecting effects, with given levels of significance and power, becomes much larger than that in a randomized controlled trial. Alternatively, the power to detect an effect with a given sample size becomes low.

Brion et al. (2013) show that the power for the Mendelian randomization study with one instrument (one SNP or multiple SNP predictor) can be expressed as a function of the non-centrality parameter (NCP) of the test for the 2SLS regression coefficient to be zero. They further show that the NCP depends on a number of parameters: the causal effect of X on Y, the OLS estimate of it, the proportion of the exposure variable explained by the genetic predictor, and the variances of the X and Y variables. They provide an online tool to perform such power calculations, which we have utilized.\footnote{The web tool is available at http://glimmer.rstudio.com/kn3in/mRnd/} Using the OLS and IV estimates obtained above, the observed variances of the X and Y variables, and the explanatory power of a genetic instrument in our data (0.048, i.e. ~5%), the required sample size for a test of no causal relationship with 0.05 type I error and 0.7 power is approximately N=13500. Therefore, our result of no causal effect should be interpreted with caution, as it may reflect our inability to reject the null hypothesis due to the relatively small sample size.\footnote{There is an earlier study that uses IV design with the same data that reports statistically significant effects of obesity on labor market outcomes (Böckerman et al., 2016). This suggests that the sample size is not an important issue \textit{per se}.} On the other hand, the nature of our outcome measure, register-based long-term earnings, should substantially decrease the measurement error and thereby improve the precision of estimation compared to e.g., Tyrrell et al. (2016).

Table 4 contains additional estimation results after controlling for other genetic risk scores and observable differences in cognitive skills and personality characteristics. Allen
et al. (2010) identified five loci for which the height-associated variant was also correlated with variants associated with other traits and diseases, particularly bone mineral density, rheumatoid arthritis, type 1 diabetes, psoriasis and obesity. Although the use of the genetic risk score significantly reduces the risk that an individual gene variant will bias the results (Palmer et al., 2012), we also added the genetic risk scores for blood pressure, total cholesterol, and triglycerides to control for potential pleiotropy. The results remain intact after this adjustment.

The most interesting finding from Table 4 is that the relationship between height and labor market outcomes is not statistically significant in the OLS estimation after adding the measures for cognitive skills to the set of covariates. Although the decline in the point estimate is partly attributed to the reduction in the sample size, this result suggests that physical height may be a marker of beneficial circumstances for developing higher cognitive skills (cf. Case and Paxson, 2008). However, the measures of cognitive skills are only available for a subsample of the original YFS data, implying that we have to consider these results with caution. Accounting for personality characteristics does not significantly impact on the estimates of how height influences earnings and employment.16

The YFS dataset is quite small for estimating separate results for subsamples, but we estimated models by gender to identify potential differences. The OLS estimates for both men and women are almost identical (not reported). The IV estimates are not statistically significant. In particular, the estimate for men is also very close to zero. There is earlier empirical evidence from other countries showing that the height premium is larger for

16 Hübler (2013) has shown using German panel data that the height premium disappears after controlling for personality traits because tall persons are relatively risk tolerant compared to their shorter peers.
men (Hübler, 2016). Using a self-reported categorical annual household income from a single year and a genetic instrument, Tyrrell et al. (2016) report that the effect of height on annual household income is approximately twice as strong in men as in women. The previous Finnish OLS estimates also give larger effects for men, but they are available from data that contain a different age cohort (Böckerman et al., 2010; Böckerman and Vainiomäki, 2013). The institutional setting of the labor market is also potentially important. In particular, wage compression and women’s high labor force participation rate in Finland in the six age cohorts that we examine may contribute to the small difference in the height premium between women and men.

We additionally estimated separate models for the recession years (2009-2012) and pre-recession period (2001-2008) to account for the possible business cycle variation in the height effects, as reported in Böckerman and Vainiomäki (2013). Our OLS estimates (not reported) show that the point estimate for the height effect almost doubles in the recession compared to the pre-recession period, but the IV estimates remain insignificant in both periods.

Our baseline IV estimates used non-weighted genetic score based on 180 SNPs. As a final robustness check, we use alternative genetic scores. First, the IV estimates based on the weighted genetic score based on 180 SNPs reported in Appendix A4 are similar compared to the ones that used non-weighted genetic score in Table 2. Second, there is empirical literature that identifies more relevant SNPs for height (Wood et al., 2014) compared with the 10% explained by the 180 SNPs included in the genetic score used in the baseline estimates of this study. A more recent GWAS identified altogether 697 SNPs at a genome-wide significant level that explained 16% of phenotypic variance. The effect sizes of the weighted height scores for these are 1.97 cm/SD(GRS) for 180 SNPs and...
2.80 cm/SD(GRS) for 697 SNPs in the entire YFS population. For this reason, we have also used information on 697 SNPs to construct the genetic score.\textsuperscript{17} Genotype imputation was performed using SHAPEIT v1 (Delaneau, 2011) and IMPUTE2 (Howie, 2009) software and the 1000G Phase I Integrated Release Version 3 as a reference panel (1000 Genomes Project Consortium, 2010). The results using non-weighted and weighted score based on 697 SNPs are reported in Appendix A5. The IV estimates using the alternative genetic score are similar to our baseline estimates in Table 2.

\section*{Discussion}

Using the genetic score as an instrument for measured height, we find that the IV point estimate for the height premium is lower than the OLS estimate and is not significantly different from zero. Taken at face value, this suggests that the OLS estimates for the quantitative size of the height premium in the literature may be upward biased. The use of the genetic score as an instrument for height accounts for potential confounders that are related to socioeconomic background, early life conditions and parental investments that have not been systematically accounted for in previously reported empirical literature.

\textsuperscript{17} The use of the 180 SNPs is not an important limitation \textit{per se} because the most important SNPs with largest effect sizes are usually identified first. This implies that the use of additional SNPs would most likely not improve the power of the first stage of the IV regressions much. Because our data is relatively small additional SNPs with smaller effect size can even produce results that have the wrong signs compared to GWAS. A large set of genetic instrumental variables would also increase the possibility that some of the instruments are invalid due to pleiotropy (Bowden et al., 2015).
As the genetic score (used as an instrument for height) was not associated with earnings or employment while true height was, the associations between height and labor market outcomes are not causal effects. Instead, our results suggest that physical height is a marker of beneficial circumstances for developing higher cognitive skills during childhood or adolescence due to family background. The disappearance of the height effect by inclusion of cognitive skills as an explanatory variable supports this interpretation.

Our results are broadly consistent with those reported in a twin design for monozygotic (identical) twins by Böckerman and Vainiomäki (2013). The authors found an insignificant height premium for men after controlling for genetic differences between twins, which supports that the cross-sectional OLS results are driven by unobserved differences such as cognitive skills. For women, they found that the height premium prevails in earned income, but not in capital income, for identical twins. This suggests that discrimination is a potential explanation.

Our study has two limitations. First, the number of observations in the baseline estimations is ~2,000. A larger data set would be needed to provide more power in order to obtain more tightly estimated effects. Second, the Young Finns Study that we use in the estimations is not nationally representative with respect to the total population in Finland, which consists also of the older age cohorts. The data are representative only for the selected six age cohorts in 1980. Further studies are needed to confirm that the patterns prevail in all age cohorts and other institutional settings.
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Figure 1. Illustration of Mendelian randomization.

\[ \beta_{IV} = \frac{\beta_{GY}}{\beta_{GX}} = \frac{\text{cov}(G, Y)}{\text{cov}(G, X)} \]

\[ \beta_{OLS} = \frac{\text{cov}(X, Y)}{\text{var}(X)} \]

\[ E[\hat{\beta}_{OLS}] = \beta_{XY} + \beta_{XZ} \cdot \beta_{ZY} \]

\[ E[\hat{\beta}_{IV}] = \beta_{XY} \] true causal effect

Note: formulas assume one X, G and Z variable.
Appendix A2. Kernel density plot of the genetic score.
Table 1. Sample characteristics.

|                              | Height below gender specific mean | Height above gender specific mean | Difference | t-statistics | N  |
|------------------------------|----------------------------------|----------------------------------|------------|--------------|----|
| Average earnings 2001-2012 (euros per year) | 23391.17 (13185.86) | 25663.80 (16405.44) | -2272.627 | -3.399*** | 1982 |
| Share of years employed, 2001-2012 | 0.858 (0.240) | 0.878 (0.211) | -0.020 | -1.979** | 1982 |
| Genetic score for height | 177.836 (0.268) | 181.816 (0.269) | -3.980 | -10.474*** | 1982 |
| Married (2001) | 0.438 (0.496) | 0.458 (0.498) | -0.020 | -0.903 | 1982 |

**Cognitive skills (2011)**

|                              | Mean (SD) | Mean (SD) | t-statistics | N  |
|------------------------------|-----------|-----------|--------------|----|
| Overall cognitive performance | -0.029 (1.012) | 0.034 (0.973) | -0.063 | -1.142 | 1302 |
| Visual and episodic memory and visuospatial associative learning | 0.001 (1.013) | 0.019 (0.951) | -0.018 | -0.335 | 1302 |
| Reaction time | -0.019 (0.995) | 0.064 (0.994) | -0.083 | -1.511 | 1302 |
| Rapid visual information processing | -0.013 (1.003) | 0.087 (0.964) | -0.100 | -1.829* | 1302 |
| Spatial working memory | -0.014 (0.970) | 0.003 (0.980) | -0.017 | -0.315 | 1302 |

**Personality characteristics (1983)**

|                              | Mean (SD) | Mean (SD) | t-statistics | N  |
|------------------------------|-----------|-----------|--------------|----|
| Aggression | 3.850 (0.912) | 3.806 (0.876) | 0.044 | 0.803 | 1042 |
| Leadership | 4.221 (0.910) | 4.322 (0.862) | -0.100 | -1.830* | 1042 |
| Responsibility | 4.861 (1.027) | 4.846 (1.075) | 0.015 | 0.227 | 1042 |
| Eagerness-Energy | 4.686 (0.705) | 4.757 (0.706) | -0.071 | -1.629 | 1042 |

**Family background (1980)**

|                              | Mean (SD) | Mean (SD) | t-statistics | N  |
|------------------------------|-----------|-----------|--------------|----|
| Annual income, mother (euros) | 4497.34 (3519.50) | 4771.26 (3481.42) | -273.91 | -1.691* | 1868 |
| Annual income, father (euros) | 8525.22 (5061.18) | 9010.45 (6437.12) | -485.23 | -1.811* | 1868 |
| High education, mother | 0.068 (0.251) | 0.078 (0.268) | -0.010 | -0.865 | 1982 |
| High education, father | 0.094 (0.292) | 0.111 (0.314) | -0.017 | -1.259 | 1982 |

**Other genetic risk scores**

|                              | Mean (SD) | Mean (SD) | t-statistics | N  |
|------------------------------|-----------|-----------|--------------|----|
| Blood pressure | 30.638 | 30.267 | 0.371 | 2.568** | 1957 |
|                | (3.139) | (3.247) |     |     |     |
|----------------|---------|---------|-----|-----|-----|
| Total cholesterol | 27.460  | 27.414  | 0.046 | 0.327 | 1957 |
|                | (3.123) | (3.057) |     |     |     |
| Triglycerides  | 26.301  | 25.958  | 0.343 | 2.634*** | 1957 |
|                | (2.860) | (2.900) |     |     |     |
Table 2. Baseline results.

|                                | Panel A: OLS | Panel B: IV |
|--------------------------------|--------------|-------------|
|                                | Coefficient  | t-statistics| Coefficient  | t-statistics |
|                                | (95% CI)     |             | (95% CI)     |             |
| **Log of average earnings, 2001-2012** |              |             |              |             |
| Height                         | 0.013        | 3.91        | 0.009        | 0.84        |
|                                | (0.006, 0.020)***** |             | (-0.012, 0.030) |             |
| F-statistics                   | ..           | 187.71      |              |             |
| Mean outcome                   | 9.867        |             | 9.867        |             |
| N                              | 1982         |             | 1982         |             |
| **Share of years employed, 2001-2012** |              |             |              |             |
| Height                         | 0.003        | 3.12        | 0.001        | 0.33        |
|                                | (0.001, 0.004)***** |             | (-0.004, 0.006) |             |
| F-statistics                   | ..           | 187.71      |              |             |
| Mean outcome                   | 0.868        |             | 0.868        |             |
| N                              | 1982         |             | 1982         |             |

Notes: Earnings are measured as the log of average earnings over the period of 2001-2012. Employment is measured as the average share of employment years over the period of 2001-2012. The mean values for the dependent variables are reported. Height was measured in 2001. All models include controls for the birth month, birth year effects, gender and parental education (1980). The instrument used in the IV models is the genetic score for height based on genetic markers. Angrist-Pischke multivariate F-tests of the excluded instrument are reported for the IV models. 95% confidence intervals based on heteroscedasticity-robust standard errors are reported in parentheses: significant at *10% **5% and ***1% levels.
Table 3. Comparison of observables by instrument value.

|                                | Below gender specific average height score | Above gender specific average height score | Difference | t-statistics | N  |
|--------------------------------|------------------------------------------|-------------------------------------------|------------|--------------|----|
| Average earnings 2001-2012 (euros per year) | 24389.03 (14468.18)                      | 24661.27 (15354.84)                      | -272.239   | -0.406       | 1982 |
| Share of years employed 2001-2012 | 0.862 (0.239)                            | 0.875 (0.213)                            | -0.013     | -1.306       | 1982 |
| Height in cm (2001)             | 170.684 (8.901)                          | 173.661 (9.061)                          | -2.977     | -7.375***    | 1982 |
| Married (2001)                  | 0.436 (0.496)                            | 0.459 (0.499)                            | -0.023     | -1.028       | 1982 |
| Cognitive skills (2011)         |                                          |                                           |            |              |     |
| Overall cognitive performance  | -0.029 (1.014)                           | 0.034 (0.972)                            | -0.064     | -1.154       | 1302 |
| Visual and episodic memory and visuospatial associative learning | -0.034 (1.011) | 0.053 (0.951) | -0.087 | -1.606 | 1302 |
| Reaction time                   | 0.001 (0.999)                            | 0.045 (0.992)                            | -0.043     | -0.787       | 1302 |
| Rapid visual information processing | 0.049 (0.976) | 0.027 (0.993) | 0.022 | 0.395 | 1302 |
| Spatial working memory         | -0.017 (0.977)                           | 0.007 (0.974)                            | -0.025     | -0.453       | 1302 |
| Personality characteristics (1983) |                                          |                                           |            |              |     |
| Aggression                      | 3.826 (0.900)                            | 3.830 (0.889)                            | -0.004     | -0.075       | 1042 |
| Leadership                      | 4.253 (0.914)                            | 4.289 (0.862)                            | -0.036     | -0.655       | 1042 |
| Responsibility                  | 4.826 (1.060)                            | 4.879 (1.043)                            | -0.053     | -0.809       | 1042 |
| Eagerness-Energy                | 4.733 (0.712)                            | 4.711 (0.701)                            | 0.022      | 0.500        | 1042 |
| Family background (1980)        |                                          |                                           |            |              |     |
| Annual income, mother (euros)   | 4599.33 (3601.71)                        | 4667.80 (3404.99)                       | -68.467    | -0.422       | 1868 |
| Annual income, father (euros)   | 8548.13 (5873.54)                        | 8979.62 (5709.11)                       | -431.495   | -1.610       | 1868 |
| High education, mother          | 0.061 (0.239)                            | 0.084 (0.278)                            | -0.024     | -2.043**     | 1982 |
| High education, father          | 0.099 (0.298)                            | 0.106 (0.308)                            | -0.008     | -0.557       | 1982 |
| Other genetic risk scores       |                                          |                                           |            |              |     |
| Blood pressure                  | 30.537 (3.192)                           | 30.371 (3.203)                           | 0.165      | 1.144        | 1957 |
| Total cholesterol               | 27.498 (3.137)                           | 27.379 (3.043)                           | 0.119      | 0.849        | 1957 |
| Triglycerides                   | 26.168 (2.879)                           | 26.094 (2.892)                           | 0.074      | 0.570        | 1957 |
Notes: Significant at *10% **5% and ***1% levels.

Table 4. Additional results, controlling for cognitive skills and personality characteristics.

|                         | Panel A: OLS                        | Panel B: IV                        |
|-------------------------|-------------------------------------|------------------------------------|
|                         | Adding other genetic risk scores as a control | Adding cognitive skills as a control | Adding personality characteristics as a control | Adding other genetic risk scores as a control | Adding cognitive skills as a control | Adding personality characteristics as a control |
| Log of average earnings, 2001-2012 |                          |                                    |                                |                                      |                                    |                                      |
| Height                  | 0.013 (0.007, 0.020)*** | 0.004 (-0.004, 0.011)               | 0.014 (0.005, 0.023)***       | 0.009 (-0.012, 0.030)               | 0.007 (-0.017, 0.030)               | 0.007 (-0.019, 0.033)               |
| F-statistics            | . . 180.54                  | 135.69                              | 114.97                         |
| N                       | 1957                       | 1302                                | 1042                           |
| Share of years employed, 2001-2012 |                  |                                    |                                |                                      |                                    |                                      |
| Height                  | 0.003 (0.001, 0.005)*** | 0.000 (-0.002, 0.002)               | 0.003 (0.001, 0.006)***       | 0.001 (-0.004, 0.007)               | 0.000 (-0.006, 0.006)               | 0.001 (-0.006, 0.007)               |
| F-statistics            | . . 180.54                  | 135.69                              | 114.97                         |
| N                       | 1957                       | 1302                                | 1042                           |

Notes: The effects of additional controls are not reported. Earnings are measured as the log of average earnings over the period of 2001-2012. Employment is measured as the average share of employment years over the period of 2001-2012. Height was measured in 2001. All models include controls for the birth month, birth year effects, gender and parental education (1980). The instrument used in the IV models is the genetic score for height based on genetic markers. Angrist-Pischke multivariate F-tests of the excluded instrument are reported for the IV models. 95% confidence intervals based on heteroscedasticity-robust standard errors are reported in parentheses: Significant at *10% **5% and ***1% levels.
Appendix A1. SNPs and the genotyped risk alleles used for the genetic score of height.

| SNP id    | Effect allele | Beta  | Imputation quality (info) |
|-----------|---------------|-------|---------------------------|
| rs425277  | T             | 0.022 | 0.996                     |
| rs2284746 | C             | -0.04 | 0.989                     |
| rs1738475 | C             | 0.025 | 0.9987                    |
| rs4601530 | T             | -0.028| 1                         |
| rs7532866 | A             | 0.021 | 0.9793                    |
| rs2154319 | T             | -0.03 | 0.9161                    |
| rs17391694| T             | 0.042 | 0.995                     |
| rs6699417 | T             | 0.021 | 1                         |
| rs10874746| T             | -0.024| 0.9969                    |
| rs9428104 | A             | -0.041| 0.9886                    |
| rs11205277| A             | -0.046| 0.9939                    |
| rs17346452| T             | -0.04 | 0.9861                    |
| rs1325598 | A             | -0.022| 0.9968                    |
| rs1046934 | A             | -0.044| 1                         |
| rs10863936| A             | -0.021| 1                         |
| rs6684205 | A             | -0.028| 0.9937                    |
| rs1118346 | T             | -0.025| 0.9893                    |
| rs10799445| A             | 0.032 | 0.9863                    |
| rs4665736 | T             | 0.029 | 0.9337                    |
| rs6714546 | A             | -0.026| 0.9911                    |
| rs17511102| A             | -0.06 | 0.9565                    |
| rs2341459 | T             | 0.025 | 0.9998                    |
| rs12474201| A             | 0.028 | 0.9939                    |
| rs3791675 | T             | -0.053| 0.9959                    |
| rs11684404| T             | -0.028| 0.9987                    |
| rs7567288 | T             | -0.032| 0.9832                    |
| rs7567851 | C             | 0.037 | 1                         |
| rs1351164 | T             | 0.034 | 0.995                     |
| rs12470505| T             | 0.041 | 0.9975                    |
| rs2629046 | T             | 0.024 | 0.9991                    |
| rs2580816 | T             | -0.045| 0.9983                    |
| rs12694997| A             | -0.024| 0.9997                    |
| rs2597513 | T             | -0.036| 0.9998                    |
| rs13088462| T             | -0.052| 0.9976                    |
| rs2336725 | T             | -0.027| 0.9906                    |
| rs9835332 | C             | -0.026| 1                         |
| rs17806888| T             | 0.036 | 0.9641                    |
| rs9863706 | T             | -0.031| 0.9944                    |
| rs6439167 | T             | -0.034| 0.9978                    |
| rs9844666 | A             | -0.024| 0.9985                    |
| rs724016  | A             | -0.07 | 0.9984                    |
| rs572169  | T             | 0.033 | 0.9989                    |
| rs720390 |  A  |  0.029 | 0.9953 |
| rs2247341 |  A  |  0.025 | 0.9911 |
| rs6449353 |  T  |  0.075 | 0.9553 |
| rs17081935 |  T  |  0.03  | 0.9986 |
| rs7697556 |  T  |  0.028 | 0.9835 |
| rs788867 |  T  | -0.043 | 0.987  |
| rs10010325 |  A  |  0.024 | 0.9998 |
| rs6449353 |  T  | -0.073 | 0.9995 |
| rs955748 |  A  | -0.023 | 0.994  |
| rs1173727 |  T  |  0.034 | 0.999  |
| rs11958779 |  A  | -0.027 | 0.9997 |
| rs10037512 |  T  |  0.032 | 0.9976 |
| rs13177718 |  T  | -0.04  | 0.9998 |
| rs1582931 |  A  | -0.023 | 0.8666 |
| rs274546 |  A  | -0.029 | 0.9986 |
| rs526896 |  T  |  0.03  | 0.9177 |
| rs4282339 |  A  | -0.036 | 0.9956 |
| rs12153391 |  A  | -0.03  | 0.9429 |
| rs889014 |  T  | -0.03  | 0.9858 |
| rs422421 |  T  | -0.031 | 0.9362 |
| rs6879260 |  T  | -0.022 | 0.9283 |
| rs3812163 |  A  | -0.036 | 0.9722 |
| rs1047014 |  T  | -0.032 | 0.996  |
| rs806794 |  A  |  0.052 | 0.9911 |
| rs3129109 |  T  | -0.032 | 0.9701 |
| rs2256183 |  A  |  0.04  | 0.9995 |
| rs6457620 |  C  | -0.029 | 0.9991 |
| rs2780226 |  T  | -0.076 | 0.9876 |
| rs6457821 |  A  | -0.104 | 0.9181 |
| rs9472414 |  A  | -0.026 | 0.9999 |
| rs9360921 |  T  | -0.042 | 0.9999 |
| rs310405 |  A  |  0.026 | 0.9897 |
| rs7759938 |  T  | -0.045 | 0.9878 |
| rs1046943 |  A  |  0.02  | 0.9972 |
| rs961764 |  C  | -0.024 | 0.9725 |
| rs1490384 |  T  |  0.034 | 0.9748 |
| rs6569648 |  T  | -0.04  | 0.9992 |
| rs7763064 |  A  | -0.048 | 0.9938 |
| rs543650 |  T  | -0.034 | 0.9978 |
| rs9456307 |  A  | -0.048 | 0.8046 |
| rs798489 |  T  | -0.048 | 0.9921 |
| rs4470914 |  T  |  0.029 | 0.9998 |
| rs12534093 |  A  | -0.034 | 0.9638 |
| rs1708299 |  A  |  0.04  | 0.9873 |
| rs6959212 |  T  | -0.024 | 0.9834 |
| rs42235 |  T  |  0.057 | 0.966  |
| rs822552 |  C  | -0.025 | 0.8303 |
| rs          | Allele | P-Value | OR  |
|------------|--------|---------|-----|
| rs2110001  | C      | 0.031   | 0.9028 |
| rs1013209  | T      | 0.025   | 0.9704 |
| rs7460090  | T      | 0.058   | 0.993 |
| rs6473015  | A      | -0.029  | 0.9998 |
| rs6470764  | T      | -0.05   | 0.9929 |
| rs12680655 | C      | 0.028   | 0.9998 |
| rs7864648  | T      | 0.022   | 0.9983 |
| rs11144688 | A      | -0.049  | 0.4495 |
| rs7853377  | A      | -0.024  | 0.9623 |
| rs8181166  | C      | 0.026   | 0.9772 |
| rs2778031  | T      | 0.031   | 0.9584 |
| rs9969804  | A      | 0.03    | 1 |
| rs1257763  | A      | 0.069   | 0.5821 |
| rs473902   | T      | 0.065   | 0.5173 |
| rs7027110  | A      | 0.031   | 1 |
| rs1468758  | T      | -0.026  | 0.9938 |
| rs751543   | T      | 0.026   | 0.8534 |
| rs7466269  | A      | 0.032   | 0.9997 |
| rs7849585  | T      | 0.029   | 0.9933 |
| rs7909670  | T      | -0.021  | 0.9985 |
| rs2145998  | A      | -0.026  | 0.9997 |
| rs11599750 | T      | -0.028  | 0.9997 |
| rs2237886  | T      | 0.046   | 0.9947 |
| rs7926971  | A      | -0.023  | 0.9948 |
| rs1330     | T      | 0.022   | 0.9995 |
| rs10838801 | A      | -0.027  | 0.9935 |
| rs1814175  | T      | 0.022   | 0.9703 |
| rs5017948  | A      | 0.027   | 0.8788 |
| rs3782089  | T      | -0.058  | 0.9562 |
| rs7112925  | T      | -0.023  | 0.9955 |
| rs634552   | T      | 0.039   | 0.9934 |
| rs494459   | T      | 0.02    | 1 |
| rs654723   | A      | 0.025   | 0.8502 |
| rs2856321  | A      | -0.029  | 0.9993 |
| rs10770705 | A      | 0.033   | 0.9998 |
| rs2638953  | C      | 0.032   | 1 |
| rs2066807  | C      | -0.054  | 0.8978 |
| rs1351394  | T      | 0.06    | 0.9851 |
| rs10748128 | T      | 0.038   | 0.9985 |
| rs11107116 | T      | 0.052   | 1 |
| rs7971536  | A      | -0.028  | 0.924 |
| rs11830103 | A      | -0.035  | 0.9901 |
| rs7332115  | T      | -0.023  | 1 |
| rs3118905  | A      | -0.056  | 0.9996 |
| rs7319045  | A      | 0.025   | 0.9681 |
| rs1950500  | T      | 0.034   | 0.9999 |
| rs2093210  | T      | -0.032  | 0.9512 |
| SNP          | Allele | MAF  | P-value  |
|--------------|--------|------|----------|
| rs1570106    | T      | -0.026 | 0.9985   |
| rs862034     | A      | -0.028 | 0.9888   |
| rs7155279    | T      | -0.024 | 0.9621   |
| rs16964211   | A      | -0.05  | 0.9998   |
| rs7178424    | T      | -0.021 | 0.9999   |
| rs10152591   | A      | 0.041  | 0.9997   |
| rs12902421   | T      | -0.062 | 0.962    |
| rs5742915    | T      | -0.031 | 0.9325   |
| rs11259936   | A      | -0.044 | 1        |
| rs16942341   | T      | -0.13  | 0.968    |
| rs2871865    | C      | 0.057  | 0.9402   |
| rs4965598    | T      | -0.028 | 0.989    |
| rs11648796   | A      | -0.034 | 0.9535   |
| rs26868      | A      | 0.034  | 0.9899   |
| rs1659127    | A      | 0.027  | 0.9533   |
| rs8052560    | A      | 0.029  | 0.7692   |
| rs4640244    | A      | 0.024  | 0.9952   |
| rs3110496    | A      | -0.022 | 0.9939   |
| rs3764419    | A      | -0.035 | 0.9994   |
| rs17780086   | A      | 0.028  | 0.9857   |
| rs1043515    | A      | -0.023 | 0.9968   |
| rs4986172    | T      | -0.032 | 0.9763   |
| rs2072153    | C      | 0.021  | 1        |
| rs4605213    | C      | 0.021  | 0.9571   |
| rs227724     | A      | -0.03  | 0.9471   |
| rs2079795    | T      | 0.04   | 0.9995   |
| rs2665838    | C      | -0.042 | 0.9712   |
| rs11867479   | T      | 0.025  | 0.9982   |
| rs4800452    | T      | 0.051  | 0.9995   |
| rs9967417    | C      | -0.038 | 0.9695   |
| rs17782313   | T      | -0.028 | 0.9995   |
| rs12982744   | C      | -0.03  | 0.9753   |
| rs7507204    | C      | 0.036  | 0.9457   |
| rs891088     | A      | -0.029 | 1        |
| rs4702910    | C      | -0.031 | 0.7501   |
| rs2279008    | T      | 0.025  | 0.9958   |
| rs17318596   | A      | 0.032  | 0.9993   |
| rs1741344    | T      | -0.023 | 0.997    |
| rs2145272    | A      | -0.039 | 0.9976   |
| rs7274811    | T      | -0.041 | 0.9961   |
| rs143384     | A      | -0.063 | 0.8833   |
| rs237743     | A      | 0.041  | 1        |
| rs2834442    | A      | 0.026  | 0.9829   |
| rs4821083    | T      | 0.031  | 0.9997   |

Note: All SNPs are imputed with MACH software and HapMap release 22 as a reference panel.
Appendix A3. Eigenvalues and proportions of variance explained by principal components for cognitive performance used in the study.

| Component                                                | Eigenvalue | Explained variance |
|----------------------------------------------------------|------------|--------------------|
| Component for Paired Associates Learning test (PAL)      | 7.72       | 0.55               |
| Component for Spatial Working Memory test (SWM)         | 5.06       | 0.42               |
| Component for Reaction Time test (RTI)                  | 2.29       | 0.38               |
| Component for Rapid Visual Information test (RVP)       | 5.19       | 0.74               |

Notes: Principal component analyses were performed separately for all individual tests. The first components resulting from these analyses were considered to represent cognitive performance related to the particular cognitive domain.
Appendix A4. IV results; weighted genetic score based on 180 SNPs.

|                      | Coefficient (95% CI) | t-statistics |
|----------------------|----------------------|--------------|
| **Log of average earnings, 2001-2012** |                      |              |
| Height               | 0.006 (-0.013, 0.024) | 0.59         |
| F-statistics         | 212.82                |              |
| Mean outcome         | 9.867                 |              |
| N                    | 1981                  |              |
| **Share of years employed, 2001-2012** |                      |              |
| Height               | 0.000 (-0.005, 0.005) | -0.02        |
| F-statistics         | 212.82                |              |
| Mean outcome         | 0.868                 |              |
| N                    | 1981                  |              |

Notes: The effects of additional controls are not reported. Earnings are measured as the log of average earnings over the period of 2001-2012. Employment is measured as the average share of employment years over the period of 2001-2012. Height was measured in 2001. All models include controls for the birth month, birth year effects, gender and parental education (1980). The instrument used in the IV models is the genetic score for height based on genetic markers. Angrist-Pischke multivariate F-tests of the excluded instrument are reported for the IV models. 95% confidence intervals based on heteroscedasticity-robust standard errors are reported in parentheses: Significant at *10% **5% and ***1% levels.
Appendix A5. IV results; genetic score based on 697 SNPs.

|                                | Panel A: IV, non-weighted score | Panel B: IV, weighted score |
|--------------------------------|---------------------------------|-----------------------------|
|                                | Coefficient (95% CI)            |                          |
|                                | t-statistics                    | t-statistics                |
| **Log of average earnings, 2001-2012** |                                 |                             |
| Height                         | 0.006 (-0.007, 0.019)           | 0.003 (-0.009, 0.015)      |
|                                | 0.94                            | 0.50                        |
| F-statistics                   | 453.57                          | 488.91                      |
| Mean outcome                   | 9.867                           | 9.867                       |
| N                              | 1982                            | 1982                        |
| **Share of years employed, 2001-2012** |                                 |                             |
| Height                         | 0.001 (-0.003, 0.004)           | -0.000 (-0.004, 0.003)     |
|                                | 0.52                            | -0.20                       |
| F-statistics                   | 453.57                          | 488.91                      |
| Mean outcome                   | 0.868                           | 0.868                       |
| N                              | 1982                            | 1982                        |

Notes: The effects of additional controls are not reported. Earnings are measured as the log of average earnings over the period of 2001-2012. Employment is measured as the average share of employment years over the period of 2001-2012. Height was measured in 2001. All models include controls for the birth month, birth year effects, gender and parental education (1980). The instrument used in the IV models is the genetic score for height based on genetic markers. Angrist-Pischke multivariate F-tests of the excluded instrument are reported for the IV models. 95% confidence intervals based on heteroscedasticity-robust standard errors are reported in parentheses: Significant at *10% **5% and ***1% levels.