Educational attainment and drinking behaviors: Mendelian randomization study in UK Biobank

Tao Zhou¹, Dianjianyi Sun¹, Xiang Li¹, Hao Ma¹, Yoriko Heianza¹, Lu Qi¹,²
¹Department of Epidemiology, School of Public Health and Tropical Medicine, Tulane University, New Orleans, LA
²Department of Nutrition, Harvard T.H. Chan School of Public Health, Boston, MA

Abstract

Educational attainment has been associated with drinking behaviors in observation studies. We performed Mendelian randomization analysis to determine whether educational attainment causally affected drinking behaviors, including amount of alcohol intakes (in total and various types), drinking frequency, and drinking with or without meals among 334,507 White British participants from the UK Biobank cohort. We found that genetically instrumented higher education (one additional year) was significantly related to higher total amount of alcohol intake (inverse variance weighted method (IVW): beta=0.44, 95% confidence interval (CI) 0.40 to 0.49, \( P=1.57 \times 10^{-93} \)). The causal relations with total amount and frequency of alcohol drinking were more evident among women. In analyses of different types of alcohol, higher educational attainment showed the strongest causal relation with more consumption of red wine (IVW beta=0.34, 95% CI 0.32 to 0.36, \( P=2.65 \times 10^{-247} \)), followed by white wine/champagne, in a gender-specific manner. An inverse association was found for beer/cider and spirits. In addition, we found that one additional year of educational attainment was causally related to higher drinking frequency (IVW beta=0.54, 95% CI 0.51 to 0.57, \( P=4.87 \times 10^{-230} \)) and a higher likelihood to take alcohol with meals (IVW: odds ratio (OR)=3.10, 95% CI 2.93 to 3.29, \( P=0.00 \times 10^{0} \)). The results indicate causal relations of higher education with intake of more total alcohol especially red wine, and less beer/cider and spirits, more frequent drinking, and drinking with meals, suggesting the importance of improving drinking behaviors, especially among people with higher education.

Introduction

Worldwide, alcohol consumption is a leading cause of ill health and premature mortality [1]. In 2016, the harmful use of alcohol including intoxication resulted in some 3 million deaths (5.3% of all deaths) [2]. In contrast, studies suggested benefits of light-to-moderate
alcohol consumption on Coronary Heart Disease (CHD) risk and blood pressure [3–5], and evidence has also shown that consumption of certain types of alcohol, such as red wine [6], may lower the risks of premature mortality and various diseases including cardiovascular disease and diabetes [7, 8]. However, more recent researches increasingly show no protective or adverse effect of light-to-moderate alcohol on health [1, 9–13] and the potential benefits of alcohol consumption are likely to be offset when overall health risks were considered [1]. According to the Dietary Guidelines for Americans (2015–2020), individuals who do not drink alcohol are not recommended to start drinking for any reason. On the other hand, drinking habits, especially the frequency of drinking and whether drinking with meals, have been related to health events [14, 15]. Given the amount of alcohol, more frequent drinking and drinking with meals showed beneficial associations with a lower BMI and post-prandial blood pressure [16].

Epidemiological studies suggest that individuals with higher educational attainment tend to consume more alcohol, compared to those with lower educational attainment, though other studies showed the opposite trend [17–22]. Further, a previous study provided evidence that educational attainment was associated with different types of alcohol intake, leading to different health consequences [23]. However, most of the prior evidence about education and alcohol consumption comes from observational studies, which might be biased by confounding and reverse causation. Few studies have assessed the causality between education and alcohol drinking behaviors.

Mendelian randomization (MR) uses genetic markers as instrumental variables of exposure instead of exposure itself to facilitate causal inference [24]. This method is less likely to be affected by confounding or reverse causality [25]. A recent genome-wide association study (GWAS) revealed more than 1000 independent genome-wide-significant genetic variants associated with educational attainment [26]. Thus, in the current study, we used these variants as instruments to test the hypothesis that education may causally influence drinking behaviors in 334 507 individuals from the UK Biobank cohort. Given a potential sex difference in the relation between educational attainment and alcohol consumption [27], we also tested the causal relationship in men and women separately.

Methods

Study design and participants

UK Biobank is a large prospective study comprising nearly 0.5 million participants aged 37 to 73 years old at recruitment. Participants with self-reported white British were included in the study. Participants who 1) had a genetic relatedness with others; 2) without information on alcohol intake, 3) without genetic data; or 4) without information on covariates were excluded. Thus, 334 507 individuals were included in MR analyses. The UK Biobank study was approved by the North West Multi-Centre Research Ethics Committee and the Biomedical Committee of the Tulane University (New Orleans, Louisiana) Institutional Review Board. All participants provided written informed consent before participating in the UK Biobank study.
Education measurement and variants

Summary-level data of estimates and standard errors of the SNPs associated with educational attainment were recorded from the Social Science Genetic Association Consortium (SSGAC) [26]. Educational attainment was obtained by the touchscreen questionnaire at baseline. Participants were asked “Which of the following qualifications do you have? (You can select more than one)”. Eight options were provided under this question. To calculate years of education by the education qualification, we first mapped the above options onto the International Standard Classification of Education (ISCED) category and then imputed the years of education equivalent for each of the categories according to the previous study’s algorithm [26] (SupplementaryTable 1).

Townsend deprivation index, an indicator for socioeconomic status on the basis of national census data was obtained immediately preceding participation in UK Biobank. For occupation, we coded job class from low to high using the UK Biobank job code variable: [Elementary Occupations (coded as 1), Process, Plant and Machine Operatives (coded as 2), Sales and Customer Service Occupations (coded as 3), Personal Service Occupations (coded as 4), Skilled Trades Occupations (coded as 5), Administrative and Secretarial Occupations (coded as 6), Associate Professional and Technical Occupations (coded as 7), Professional Occupations (coded as 8), Managers and Senior Officials (coded as 9)]. For income, a categorical income variable was obtained from questionnaire, representing annual household income of <£18 000 (coded as 1), £18 000 to £30 999 (coded as 2), £31 000 to £51 999 (coded as 3), £52 000 to £100 000 (coded as 4), and >£100 000 (coded as 5). A higher coded number indicated a higher occupation/income level.

Drinking phenotypes

Data were obtained from the touchscreen questionnaire at the baseline appointment. Participants were asked about their current drinking status (never, previous, current, prefer not to answer). Alcohol intake frequency [UK Biobank Field Identifier (FID): 1558] was reported as “never”, “special occasions only”, “1 to 3 times per month”, “once or twice a week”, “3 or 4 times a week”, and “daily or almost daily” and coded from 0 to 5. Average weekly alcohol consumption of a range of drink types [red wine (FID: 1568), white wine/champagne (FID: 1578), spirits (FID: 1598), beer/cider (FID: 1588), and fortified wine (FID: 1608)] were obtained from the questionnaire. From these measures, ethanol intake (g/week) was calculated by the amount of each type of drink multiplied by its standard drink size and reference alcohol content.

Genetic data

Single nucleotide polymorphisms (SNP) genotyping, imputation, and quality control of the genetic data were performed by the UK biobank team. The majority of the sample was genotyped using the Affymetrix UK Biobank Axiom array and the rest of the sample genotyped using the Affymetrix UK BiLEVE Axiom array. A combined panel of the UK10K and 1000 Genomes phase 3 reference panels were used for imputation. Detailed information was described previously (http://www.ukbiobank.ac.uk/scientists-3/genetic-data/). We included 1199 independent SNPs (clumping were performed by Lee et al [26]). SNPs within 500 kb away from the lead SNP and correlated with r² > 0.1 were
removed [26, 28]) that were strongly associated with educational attainment from thus far the latest and largest GWAS study [26]. All these SNPs were in Hardy-Weinberg equilibrium (HWE) P > 1E-12 within the white British participants. Any palindromic SNPs with minor allele frequency above 0.42 was removed (Supplementary Table 2). The instruments selected from this study explained approximately 4.3% of the variance in education.

To evaluate the correlations of education genetic instruments with occupation and income, we develop a polygenic score (PGS) for educational attainment using the following formula:

\[
\text{PGS} = (\beta_1 \times \text{SNP}_1 + \beta_2 \times \text{SNP}_2 + \cdots + \beta_n \times \text{SNP}_n) \times \left(\frac{\text{number of selected SNPs}}{\sum \beta}\right),
\]

where \(\beta\) was the regression coefficient associated with SNP which obtained from the previous GWAS [26]. The PGS ranges from 1090 to 1346, with each unit corresponding to one effect allele and a higher score indicating a higher genetic predisposition to higher educational attainment.

**Sex differences assessment and sensitivity analyses**

To test the hypothesis that the effects of educational attainment on alcohol intake patterns may differ in men and women, we then performed analyses separately in each sex to examine whether the effects of educational attainment on alcohol intake may differ in women and men. We compared the \(\beta\) values for women and men by using z score method:

\[
z = \frac{\beta_1 - \beta_2}{\sqrt{\text{SE}_1^2 + \text{SE}_2^2}}.
\]

Given that smoking and socioeconomic status are factors that highly correlated with educational attainment and alcohol drinking, we performed stratified analyses to assess the influence of these factors. We measured the association between educational attainment and drinking behaviors in each level taken by the smoking status (nonsmokers and smoker), education (<16 years and ≥16 years) and income (<£31 000 and ≥£31 000). We further evaluate the role of socioeconomic status in the established association between educational attainment and drinking behaviors by adjusting variables such as income, occupation and Townsend score in the model individually. We also performed mediation analysis to estimate for the association of education on drinking behaviors explained by the socioeconomic status.

**Statistical analyses**

Linear regression was used to test the association between educational attainment related SNPs and alcohol intake frequency, amount of total and various types of alcohol intakes; logistic regression was used to test the association between educational attainment related SNPs and whether alcohol usually taken with meals. Amount of alcohol consumption was log (units+1) transformed. All these analyses included an adjustment for age, sex, year of birth (dummy variable), sex interacted with year of birth, and 10 genetic principal components (PCs).

An R package TwoSampleMR was used to conduct the two-sample Mendelian randomization [29]. Four complementary methods (inverse variance weighted method (IVW), mendelian randomization-Egger (MR-Egger) method, median method, and mode based estimation) with different assumptions were applied in this study [30]. A consistent
effect across these methods indicate that the causal inference is less likely to be a false positive [31]. MR-Egger intercept [32] and Mendelian randomization pleiotropy residual sum and outlier (MR-PRESSO) [33] were used to test for the presence of horizontal pleiotropy. Simulation extrapolation (SIMEX) correction was applied in mendelian MR-Egger analysis when regression dilution $\hat{P}$ for educational attainment instrument was less than 90% [34]. Results were reported for each additional year of education attainment. The mediation analysis was performed through observational analysis using Tyler’s SAS macro because of the lack of genetic instruments for socioeconomic status [35, 36]. The proportion mediated was defined as the ratio of the natural indirect effect to the total effect [35].

Statistical analyses were performed using SAS version 9.4 (SAS Institute Inc, Cary, NC) and R version 3.3.3 (The R Foundation: The R Project for Statistical Computing. http://www.R-project.org/). Two-sided $P < 0.05$ was used as the significance level.

**Results**

Baseline characteristics of the UK Biobank participants are shown in Table 1 and Table 2. Of the 334,507 participants included in the MR analysis, 179,669 (53.7%) were women. The mean age ± SD was 56.9±8.0 years, and mean BMI ± SD was 27.4 ± 4.7 kg/m$^2$. Participants with higher educational attainment tended to be younger, leaner, and drink more alcohol compared to those with lower educational attainment (all $P <0.001$). In addition, we found that one-unit increase in educational attainment PGS was associated with 0.014 increased job class and 0.007 increased income level (All $P<0.001$). One year increase in education was associated with 0.17 increased occupation level and 0.07 increased income level (All $P<0.001$).

**Amount of total and various types of alcohol intakes**

There was consistent evidence for an association between educational attainment (one additional year) and amount of alcohol intake (Figure 1A, IVW beta=0.44, 95% confidence interval (CI) 0.40 to 0.49, $P=1.57E-93$). Such relations appeared to be more evident among women (Figure 1A, IVW beta=0.59, 95% CI 0.54 to 0.64, $P=7.80E-109$). No evidence of pleiotropy was found according to the intercept of MR-Egger for the pooled participants (Supplementary Table 3, $P_{\text{intercept}}=0.92$). However, for women, slight pleiotropy was observed ($P_{\text{intercept}}=0.05$) (Supplementary Table 3). After removing outliers by using MR-PRESSO, no significant change of the association was found after correction for the detected horizontal pleiotropy ($P=0.94$). For men, though the IVW method and weighted median method showed a positive association, the results were not consistent for mode-based method (Figure 1A).

For various types of alcohol intake, we found consistent evidence between educational attainment and red wine and white wine/champagne intake in both women and men: one additional year of educational attainment was associated with more red wine intake (Figure 1B, IVW beta=0.34, 95% CI 0.32 to 0.36, $P=2.65E-247$) and white wine/champagne (Figure 1C. IVW beta=0.24, 95% CI 0.22 to 0.26, $P=5.58E-160$). The direction was consistent when using the median and mode-based method. No evidence of pleiotropy was found according to the intercept of MR-Egger (for red wine $P_{\text{intercept}}=0.78$, for white wine/
champagne $P_{\text{intercept}}=0.31$) (Supplementary Table 3). The selected educational attainment genetic variants showed different effects in women and men, indicating that the effect of educational attainment on red wine and white wine/champagne intake was more pronounced among men (For red wine, Figure 1B, IVW, beta of women and men, 0.24 vs 0.44, $P<0.001$; for white wine/champagne, Figure 1C, IVW, beta of women and men, 0.18 vs 0.30, $P<0.001$). Participant with higher educational attainment were likely to have less beer/cider (Figure 1E, IVW beta=−0.26, 95% CI −0.28 to −0.24, $P=4.07\cdot10^{-175}$) and spirits (Figure 1F, IVW beta=−0.16, 95% CI −0.17 to −0.14, $P=7.08\cdot10^{-77}$).

In addition, men with higher educational attainment were more likely to have less beer/cider compared with women (Figure 1E). We also found sex differences for the association between education and spirits intake: women with higher educational attainment were more likely to have less spirits compared with men (Figure 1F).

We did not find consistent evidence for the association between educational attainment and fortified wine in neither women nor men (Figure 1D).

**Alcohol intake frequency**

We found an association of educational attainment and alcohol intake frequency. More years in genetically determined education were significantly associated with an increased likelihood of alcohol intake frequency according to the IVW method (Figure 2, beta=0.54, 95% CI: 0.51 to 0.57, $P=4.87\cdot10^{-230}$). The direction of effect remained consistently significant across all four methods, although confidence intervals were wider in some methods. No evidence of pleiotropy was found according to the intercept of MR-Egger ($P_{\text{intercept}}=0.74$) (Supplementary Table 3). Additionally, the effect of educational attainment on alcohol intake frequency was more pronounced among women (IVW, beta of women and men, 0.60 vs 0.47, $P<0.001$) (Figure 2).

**Alcohol taken with meals**

We found that participants with higher educational attainment were more likely to take alcohol with meals (Figure 3, IVW, odds ratio (OR)=3.10, 95% CI: 2.93 to 3.29, $P=0.00\cdot10^{00}$). The intercept in the MR Egger analysis indicated no pleiotropy (Supplementary Table 3, $P_{\text{intercept}}=0.85$). Results were consistent when using weighted median and mode-based methods. For sex differences, both women and men with higher educational attainment were similarly more likely to drink alcohol with meals (Figure 3).

**Sensitivity analyses**

Full details of sensitivity analyses are provided in the supplementary files. Briefly, we found that the associations between educational attainment and intake of red wine, white wine/champagne, and drinking frequency were more pronounced among smokers (All $P<0.005$), whereas no significant difference was found for association between educational attainment and drinking with meals among smokers and nonsmokers (Supplementary Figure 1–3, Supplementary Table 3). The association between educational attainment and drinking was attenuated but remained significant after adjustment for socioeconomic status including income, occupation, and deprivation index (Figure 4). Further mediation
analysis showed that association of educational attainment with drinking behaviors was mediated by approximately up to two-third through income, half through occupation, and one-tenth through material deprivation (All $P < 0.0001$) (Figure 4). For results stratified by education and income, the associations are largely consistent across different income or education levels based on the inverse variance weighted method (Supplementary Figure 4–9, Supplementary Table 3).

**Discussion**

In this Mendelian randomization study using four complementary methods, we found consistent evidence that higher educational attainment was related to increased amount of alcohol consumption, especially red wine and white wine/champagne. In addition, higher educational attainment was also related to a higher frequency of alcohol drinking and drinking alcohol with meals.

Our finding that participants with a genetically-determined higher level of education had a higher amount of alcohol intake is consistent with previous observational studies [20]. Social context might be a key determinant on how much alcohol individuals consume. Our results indicated that socioeconomic status significantly mediated the association between educational attainment and drinking behaviors by approximately up to two-third. Educational attainment has a significant impact on social position and opportunities, which are closely related to drinking culture [37]. It was found that more educated individuals were not only more likely to have a higher wage to afford to drink more, but also were healthier than their less educated counterparts, who consumed less alcohol because of health problems [38, 39]. In addition, drinking habit picked up in college might be another reason. College attendance was found to increase the risk of heavy alcohol intake among white young adults [17], which might be partially explained by living away from parents, time socializing with friends, having a higher likelihood of risk taking or sensation seeking, and treating alcohol intake as pleasure and a normal part of life [40, 41]. We found that the effect of higher educational attainment on the amount of alcohol intake was more pronounced among women than men. Women with higher educational attainment were found to be more likely to go into high-end service industries that favor alcohol consumption, thus more likely to indulge in hazardous drinking [27, 39]. However, men enjoyed alcohol more than women, and work-place showed less impact on men’s drinking patterns [42].

For various types of alcohol intake, higher educational attainment was associated with higher red wine and lower beer/cider consumption in both women and men, while sex-difference was found for red wine, as well as white wine/champagne, beer/cider, and spirits. Higher education increases the likelihood of getting access to health information such as wine but not beer/cider or spirits reduced risks of diseases such as CHD, diabetes, and hypertension [6, 43, 44], leading to choosing healthier beverage type (e.g. red wine) and avoiding unhealthy drink (e.g. beer/cider and spirits) [45]. The sex-difference for the association between educational attainment and beverage preference might be partly explained by the differing effects of alcohol on men and women [42]. Although generally consuming less alcohol than men, women are more likely to suffer severe brain and other organ damage than men following heavy drinks [46]. Generally, women prefer wine, and
men prefer spirits. Such gender difference might overwhelm the effect of educational attainment on beverage preference [47].

Though participants with higher level of education had higher amount of alcohol intake, they tend to take alcohol in a more frequent manner (given the amount) than those with a low level of education. Data from National Health and Nutrition Examination Survey (NHANES I) and its follow-up surveys showed that participants with higher education (≥high school) had higher weekly number of drinks compared with those with lower education (<high school) [21]. Educational attainment is among the most commonly used measure of socioeconomic status in epidemiological studies [48] and a previous study suggests that people with higher socioeconomic status tend to drink smaller amounts more frequently [49]. Evidence suggests that the highest-quantity, lowest-frequency drinkers have the poorest diet quality while the lowest-quantity, higher-frequency drinkers have the best [50]. Similarly to the sex difference for the amount of alcohol intake, the association between educational attainment and alcohol intake frequency was more pronounced among women for their social position, physiological feature and health knowledge on drinking [37, 42].

Educational attainment may also affect the drinking habit of whether consuming alcohol with a meal through providing better health information. As observed in the current study, more educated participants were more likely to take alcohol with a meal for its benefit [16]. Drinking with meals slows alcohol absorption [14] and increases the rate of alcohol elimination [51]. For example, it was found that taking red wine together with the noon meal decreased the postprandial blood pressure of centrally obese, hypertensive individuals [16].

Harmful use of alcohol is a well-established risk factor for all-cause mortality [52]. Recent global data suggested even an alcohol consumption of <100g per week adversely affect health [1, 53]. Alcohol consumption can cause several types of cancer including esophageal, liver and breast cancer [54], and even light drinkers increased risk of oral cavity, pharynx, esophagus, and breast cancers [11–13]. Though alcohol consumption showed beneficial effect on kidney cancers and non-Hodgkin lymphoma, the National cancer institute suggests that any potential benefits of alcohol consumption on cancer are likely outweighed by the adverse effect of alcohol consumption. Observational studies have shown a beneficial effect of moderate alcohol consumption on CVD risk [7, 43, 55]. However, accumulating evidence using methodologies such as Mendelian randomization [10, 56], pooling cohort studies, and multivariable-adjusted meta-analyses [57] have challenged the concept that little-to-moderate alcohol consumption is universally related to a lower CVD risk [58]. Further, the observed beneficial effects are likely to be offset when taking other diseases into account [1]. Given the overall adverse effect of alcohol consumption on health, higher education might adversely affect health by increasing alcohol consumption. Thus, effort should be made to lower the amount of alcohol intake and promote healthy drinking behaviors among people with higher education.

The major strengths of this study include the large sample size, availability of detailed alcohol phenotypic data, and complementary MR methods used in the analysis. However, our study had several potential limitations. First, it is unlikely to avoid heterogeneity.
for potential pleiotropic effects or different mechanisms of association with the exposure by using so many SNPs as instruments. Although heterogeneity in causal estimates is of concern, provided that the pleiotropic effects of genetic variants are equally likely to be positive or negative, the overall estimate based on all the genetic variants may be unbiased [59]. Here we used a random-effects model in IVW methods considering for heterogeneity. Given that MR-Egger intercept suggested no evidence of horizontal pleiotropy and MR-PRESSO showed a consistent result after removal of pleiotropic outliers, our results were unlikely due to pleiotropy [30]. In addition, a recent study showed no genetic correlation between years of education and drinking frequency [60], suggesting these two phenotypes were unlikely to share some of their underlying genome-wide genetic architecture. Second, the association between educational attainment and drinking behaviors was attenuated but remained significant after adjustment for socioeconomic status. According to our mediation analysis, socioeconomic status was on the pathway between education and drinking behaviors; and therefore was not included in our main MR modeling because it might be causally affected by education [61, 62]. Third, because we studied only white British populations, our findings might not be generalizable to other ethnic groups, especially those with different cultures of alcohol intake, such as the Asian population. Fourth, we noted that UK Biobank is in the SSGAC. In the current study, the overlap rate was 334507/1131881=29.5%, and the calculated bias and type 1 error rate around 0.001 and 0.05, suggesting that findings were less likely to be biased by the sample overlap [63]. Finally, the alcohol intake measures used in this study were self-reported, which may lead to measurement bias.

In summary, our study indicates interesting relations between educational attainment drinking behaviors. Participants with higher education tend to drink more but also tend to maintain relative healthier drinking behaviors, compared with those less educated. Growing evidence has linked alcohol intake with adverse health outcomes. Our data highlights the importance of targeting people with higher education backgrounds in campaigns for lowering the amount of alcohol intake and improving healthy drinking behaviors.

**Supplementary Material**

Refer to Web version on PubMed Central for supplementary material.

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Figure 1.
Associations of educational attainment and amount of total and specific alcohol intake from Mendelian Randomization (MR) analyses. Panel A, the total amount of alcohol intake; panel B, red wine; panel C, white wine/champagne; panel D, fortified wine; panel E, beer/cider; panel F, spirits.
Figure 2.
Associations of educational attainment and alcohol intake frequency from Mendelian Randomization (MR) analyses.

| Method                        | Beta (95% CI)     | P value      |
|-------------------------------|-------------------|--------------|
| **All**                       |                   |              |
| Inverse variance weighted     | 0.54 (0.51, 0.57) | 4.87E-230    |
| MR Egger (SIMEX)              | 0.66 (0.56, 0.76) | 1.45E-35     |
| Robust adjusted profile score | 0.55 (0.51, 0.58) | 1.08E-232    |
| Simple mode                   | 0.64 (0.43, 0.85) | 2.37E-09     |
| Weighted mode                 | 0.46 (0.29, 0.63) | 6.14E-08     |
| Simple median                 | 0.53 (0.49, 0.57) | 6.40E-172    |
| Weighted median               | 0.49 (0.45, 0.53) | 2.01E-139    |
| **Women**                     |                   |              |
| Inverse variance weighted     | 0.60 (0.56, 0.64) | 7.29E-185    |
| MR Egger (SIMEX)              | 0.57 (0.44, 0.69) | 2.04E-18     |
| Robust adjusted profile score | 0.61 (0.57, 0.65) | 9.04E-185    |
| Simple mode                   | 0.47 (0.21, 0.73) | 3.88E-04     |
| Weighted mode                 | 0.48 (0.25, 0.72) | 4.89E-05     |
| Simple median                 | 0.57 (0.52, 0.62) | 2.93E-118    |
| Weighted median               | 0.53 (0.48, 0.58) | 2.56E-106    |
| **Men**                       |                   |              |
| Inverse variance weighted     | 0.47 (0.43, 0.51) | 3.18E-135    |
| MR Egger (SIMEX)              | 0.77 (0.65, 0.89) | 2.28E-35     |
| Robust adjusted profile score | 0.49 (0.45, 0.53) | 1.61E-140    |
| Simple mode                   | 0.56 (0.33, 0.79) | 2.78E-06     |
| Weighted mode                 | 0.54 (0.34, 0.73) | 7.11E-08     |
| Simple median                 | 0.48 (0.44, 0.53) | 8.41E-97     |
| Weighted median               | 0.47 (0.42, 0.52) | 3.50E-82     |
| Method                          | All OR (95% CI)     | P value |
|--------------------------------|---------------------|---------|
| Inverse variance weighted      | 3.10 (2.93, 3.29)   | 0.00E+00 |
| MR Egger (SIMEX)               | 3.68 (3.04, 4.46)   | 4.43E-38 |
| Robust adjusted profile score  | 3.14 (2.96, 3.34)   | 5.56E-292 |
| Simple mode                    | 2.84 (1.90, 4.23)   | 3.41E-07 |
| Weighted mode                  | 2.78 (1.81, 4.26)   | 3.14E-06 |
| Simple median                  | 3.06 (2.85, 3.29)   | 7.44E-202 |
| Weighted median                | 2.86 (2.66, 3.07)   | 1.90E-180 |
| Women                          |                     |         |
| Inverse variance weighted      | 2.98 (2.76, 3.22)   | 3.75E-166 |
| MR Egger (SIMEX)               | 4.49 (3.54, 5.69)   | 1.58E-33 |
| Robust adjusted profile score  | 3.08 (2.84, 3.34)   | 1.96E-160 |
| Simple mode                    | 3.19 (1.89, 5.40)   | 1.66E-05 |
| Weighted mode                  | 2.84 (1.76, 4.59)   | 2.09E-05 |
| Simple median                  | 3.00 (2.71, 3.32)   | 1.61E-100 |
| Weighted median                | 2.85 (2.57, 3.16)   | 1.12E-86 |
| Men                            |                     |         |
| Inverse variance weighted      | 3.20 (2.98, 3.45)   | 1.89E-216 |
| MR Egger (SIMEX)               | 3.15 (2.44, 4.07)   | 6.01E-18 |
| Robust adjusted profile score  | 3.23 (3.00, 3.48)   | 4.96E-207 |
| Simple mode                    | 2.80 (1.70, 4.60)   | 5.15E-05 |
| Weighted mode                  | 2.80 (1.72, 4.56)   | 3.91E-05 |
| Simple median                  | 3.01 (2.73, 3.31)   | 1.81E-112 |
| Weighted median                | 2.96 (2.67, 3.27)   | 1.04E-99 |

Figure 3.
Associations of educational attainment and alcohol taken with meals from Mendelian Randomization (MR) analyses.
Figure 4.
Estimates for the association of education with drinking behaviors explained by the socioeconomic status. *Beta value was estimated using inverse-variance weighted Mendelian randomization methods; †The proportion of the association of education with drinking behaviors that was mediated by socioeconomic status was estimated by dividing the indirect effect by the total effect (P<0.0001 for all the effects with adjustment for age and sex according to observational results). SES, socioeconomic status.
Table 1
Baseline character of UK biobank participants included in the mendelian randomization analysis

|                          | All           | Women         | Men           |
|--------------------------|---------------|---------------|---------------|
| N                        | 334 507       | 179 669       | 154 838       |
| Age, years               | 56.9±8.0      | 56.7±7.9      | 57.1±8.1      |
| Education years          | 14.9±5.1      | 14.5±5.1      | 15.3±5.1      |
| BMI, kg/cm²              | 27.4±4.7      | 27.0±5.1      | 27.8±4.2      |
| Total amount of alcohol intake, g/week | 120.5±144.9  | 76.5±92.4     | 171.6±174.9   |
| Red wine, glasses        | 4.0±5.6       | 3.4±4.7       | 4.5±6.3       |
| White wine/champagne, glasses | 2.7±4.7       | 3.4±4.9       | 2.0±4.4       |
| Fortified, glasses       | 0.2±1.3       | 0.3±1.3       | 0.2±1.3       |
| Beer/cider, pints        | 3.0±5.5       | 0.6±1.8       | 5.2±6.8       |
| Spirits, measures        | 1.8±5.3       | 1.5±4.0       | 2.2±6.2       |
| Alcohol intake frequency, n (%) |             |               |               |
| Daily or almost daily    | 71 746 (21.5) | 30 699 (17.1) | 41 047 (26.5) |
| Three or four times a week | 80 896 (24.2) | 38 920 (21.7) | 41 976 (27.1) |
| Once or twice a week     | 87 995 (26.3) | 47 374 (26.4) | 40 621 (26.3) |
| One to three times a month | 37 039 (11.1) | 23 514 (13.1) | 13 525 (8.7)  |
| Special occasions only   | 34 977 (10.5) | 24 918 (13.9) | 10 059 (6.5)  |
| Never                    | 21 649 (6.5)  | 14 133 (7.9)  | 7 516 (4.9)   |
| Alcohol taken with meals | 116 172 (67.9)| 68 226 (77.0) | 47 946 (58.1) |
| Smoking, n (%)           |               |               |               |
| Never                    | 182 914 (54.7)| 106 878 (59.5)| 76 036 (49.1) |
| Previous                 | 117 894 (35.2)| 57 333 (31.9) | 60 561 (39.1) |
| Current                  | 33 699 (10.1) | 15 458 (8.6)  | 18 241 (11.8) |
| Townsend deprivation index | −1.6±2.9     | −1.6±2.9      | −1.5±3.0      |
| Job class, n (%)         |               |               |               |
| Managers and Senior Officials | 35 820 (17.4) | 13 148 (12.4) | 22 672 (22.8) |
| Professional             | 48 765 (23.7) | 23 446 (22.1) | 25 319 (25.4) |
| Associate Professional and Technical | 34 488 (16.8) | 19 713 (18.6) | 14 775 (14.8) |
| Administrative and Secretarial | 32 909 (16)  | 26 645 (25.1) | 6 264 (6.3)   |
| Skilled Trades           | 15 589 (7.6)  | 1 780 (1.7)   | 13 809 (13.9) |
| Personal Service         | 11 922 (5.8)  | 9 985 (9.4)   | 1 937 (1.9)   |
| Sales and Customer Service | 7 172 (3.5)  | 5 524 (5.2)   | 1 648 (1.7)   |
| Process, Plant and Machine Operatives | 9 208 (4.5)  | 1 117 (1.1)   | 8 091 (8.1)   |
| Elementary               | 9 749 (4.7)   | 4 674 (4.4)   | 5 075 (5.1)   |
| Income, n (%)            |               |               |               |
| <£18 000                 | 62 615 (21.7) | 35 613 (23.9) | 27 002 (19.4) |
| £18 000 to £30 999       | 74 112 (25.7) | 39 712 (26.6) | 34 400 (24.7) |
| £31 000 to £51 999       | 76 439 (26.5) | 38 363 (25.7) | 38 076 (27.4) |
| £52 000 to £100 000      | 59 836 (20.7) | 28 444 (19.1) | 31 392 (22.6) |
|                | All       | Women     | Men       |
|----------------|-----------|-----------|-----------|
| >£100 000      | 15 507 (5.4) | 7 151 (4.8) | 8 356 (6.0) |

* 6 glasses in an average bottle
† 12 glasses in an average bottle
‡ 25 standard measures in a normal size bottle; spirits include drinks such as whisky, gin, rum, vodka, brandy.
§ Negative values will indicate relative affluence
Table 2
Baseline characteristics according to tertile of years of education in the observational studies

|                        | Tertile 1 (7 years) | Tertile 2 (10–15 years) | Tertile 3 (19–20 years) | P     |
|------------------------|---------------------|--------------------------|--------------------------|-------|
| N                      | 58,077              | 119,831                  | 162,597                  | <0.001|
| Age, years             | 61.4±6.1            | 56.7±8.0                 | 55.3±8.0                 | <0.001|
| Sex, female            | 31,530 (54.3)       | 73,229 (61.1)            | 79,005 (48.6)            | <0.001|
| BMI, kg/cm²            | 28.4±4.9            | 27.4±4.8                 | 27.0±4.6                 | <0.001|
| Total amount of alcohol intake, g/week | 109.5±155.3        | 115.6±144.7              | 128.0±140.6              | <0.001|
| Red wine, glasses *    | 2.7±4.8             | 3.8±5.6                  | 4.4±5.8                  | <0.001|
| White wine/champagne, glasses * | 1.9±4.1      | 2.8±4.8                  | 2.9±4.8                  | <0.001|
| Fortified, glasses †    | 0.3±1.6             | 0.2±1.3                  | 0.2±1.2                  | <0.001|
| Beer/cider, pints      | 4.4±7.0             | 2.8±5.5                  | 2.7±4.9                  | <0.001|
| Alcohol intake frequency, n (%) | 13,710 (48.6) | 41,017 (68.2)            | 63,912 (74.2)            | <0.001|
| Smoking, n (%)         |                     |                          |                          | <0.001|
| Never                  | 25,265 (43.8)       | 65,962 (55.2)            | 94,502 (58.2)            |       |
| Previous               | 24,035 (41.7)       | 41,650 (34.8)            | 53,956 (33.3)            |       |
| Current                | 8,332 (14.5)        | 11,937 (10.0)            | 13,792 (8.5)             |       |
| Townsend deprivation index § | −0.6±3.3         | −1.8±2.8                 | −1.8±2.8                 | <0.001|
| Job class, n (%)       |                     |                          |                          | <0.001|
| Managers and Senior Officials | 2,608 (10.7) | 12,548 (17.2)            | 21,501 (19.1)            |       |
| Professional           | 452 (1.9)           | 7,959 (10.9)             | 41,638 (36.9)            |       |
| Associate Professional and Technical | 1,445 (5.9) | 14,715 (20.2)            | 19,218 (17)              |       |
| Administrative and Secretarial | 4,283 (17.6) | 18,720 (25.7)            | 10,894 (9.7)             |       |
| Skilled Trades         | 3,770 (15.5)        | 4,958 (6.8)              | 6,999 (6.2)              |       |
| Personal Service       | 1,981 (8.1)         | 4,338 (6.0)              | 5,799 (5.1)              |       |
| Sales and Customer Service | 2,385 (9.8)   | 3,173 (4.4)              | 1,737 (1.5)              |       |
| Process, Plant and Machine Operatives | 3,312 (13.6) | 3,108 (4.3)              | 2,847 (2.5)              |       |
| Elementary             | 4,120 (16.9)        | 3,391 (4.7)              | 2,235 (2.0)              |       |
| Income, n (%)          |                     |                          |                          | <0.001|
| <£18,000               | 23,383 (53.4)       | 22,078 (21.5)            | 18,654 (12.5)            |       |
| £18,000 to £30,999     | 12,480 (28.5)       | 30,688 (29.8)            | 32,657 (21.9)            |       |
| £31,000 to £51,999     | 5,827 (13.3)        | 29,073 (28.3)            | 43,491 (29.2)            |       |
| £52,000 to £100,000    | 1,834 (4.2)         | 17,782 (17.3)            | 41,740 (28.0)            |       |
| Tertile 1 (7 years) | Tertile 2 (10–15 years) | Tertile 3 (19–20 years) | P |
|--------------------|-------------------------|-------------------------|---|
| >£100 000          | 288 (0.7)               | 3286 (3.2)              | 12 298 (8.3) |

*6 glasses in an average bottle
†12 glasses in an average bottle
‡25 standard measures in a normal size bottle; spirits include drinks such as whisky, gin, rum, vodka, and brandy
§Negative values will indicate relative affluence