INTRODUCTION

The catchability coefficient may be defined as the proportion of a fish population caught by one unit of fishing effort. Catchability (q) is hence represented by the ratio between catch rates (C/E) and stock abundance (N):

\[ q = \frac{C}{EN} \]

or equivalently by the ratio between fishing mortality (F) and fishing effort:

\[ q = \frac{F}{E} \]

The catchability coefficient has repeatedly been reported as an important source of uncertainty inherent to stock assessment models requiring fishing effort as input data (Cook and Armstrong, 1985; Pope and Shepherd, 1985; Megrey, 1989; Arreguin-Sanchez, 1996). Evidence is accumulating that the catchability of commercial fisheries is prone to non-
random variations over time (Atran and Loesch, 1995; Pascooe and Robinson, 1996) and a number of formulations have been suggested to describe these trends. These formulations relate catchability variations to the distribution of the fish population relative to: (i) the distribution of fishing (or survey) effort (Crecco and Overholtz 1990; Swain et al., 1994); (ii) the effectiveness of the vessels in finding fish (fishing power) (Gascuel et al., 1993; Millischer et al., 1999); and (iii) having found fish, actually catching them (selectivity) (Arreguin-Sanchez and Pitcher, 1999). However, the lack of data relevant to catchability dynamics has hampered the inclusion of catchability models in stock assessment procedures (Megrey, 1989). Some stock assessment models allow for the direct estimation of annual catchabilities (e.g. Fournier and Archibald, 1982) or systematic trends in catchability (e.g. the “Hybrid method”; Pope and Shepherd, 1985) from catch and effort data. However, the estimates derived from the algorithm developed by Fournier and Archibald are prone to instability, as a result of over-parameterisation (Megrey, 1989), while fitting a time trend in catchability, using the “Hybrid” method, increases the sensitivity of the fishing mortality (F) estimates to the choice of the terminal F value (Pope and Shepherd, 1985). On the other hand, other stock assessment procedures ignore time dynamics in catchability (e.g. Paloheimo, 1980; Deriso et al., 1985). The XSA (eXtended Survivors Analysis) (Shepherd, 1999) incorporates some mechanisms of density-dependent catchability, but does not account for trends related to fishing power development. ICES stock assessments are frequently carried out using XSA. In these assessments, the possible effects of changes in catchability are alleviated by reducing the length of the time series of catch per unit effort (i.e. the “tuning window”) used in the calibration equations (ICES, 2000). However, this approach is subject to criticisms, since the size of the tuning window is generally chosen arbitrarily and this procedure also means throwing away potentially useful information.

Results from past experiences and numeric simulations suggest that mis-specifying catchability could contribute to flawed stock assessments (Sampson, 1993; Chen and Paloheimo, 1998) and stock collapses (Walters and Maguire, 1996).

This study aims to examine the extent to which incorporating knowledge on annual catchability variations could contribute to enhancing the quality of stock assessments derived from XSA, which is now a usual procedure implemented by the ICES (International Council for the Exploration of the Sea) to assess fish stocks in the North-east Atlantic (e.g. ICES, 2000). A preliminary “Hybrid” analysis is carried out in order to identify the tuning fleets, which are prone to trends in catchability. XSA assessments are then carried out for a range of catchability trends, including values derived from the “Hybrid” method. The performance of these assessments is examined by means of a number of criteria based on XSA diagnostics. The above methodology is applied to the assessment of the North Sea stocks of cod, saithe, plaice and sole.

MATERIAL AND METHODS

Stock assessments

An exploratory analysis is carried out, for each tuning fleet individually, using the “Hybrid method” (Pope and Shepherd, 1985), so as to identify the tuning fleets that are subject to catchability trends. Stock assessments are then operated using the standard multi-fleet XSA (eXtended Standard Analysis) developed by Shepherd (1999). Information on the tuning fleets used in the XSA is given as annual series of landings at age and fishing effort. Fishing effort is traditionally expressed in time units fishing, possibly standardised with some measures of vessel attributes including horsepower, length or gross tonnage. The partial fishing mortality of species s and recruited age group a, harvested by fleet f, in year y, may be related to catchability q and fishing effort E, using Equation (1)

\[ F(s,f,y,a) = q(s,f,a)E(f,y) \]  

We now make two assumptions. First, we assume that the catch data used as inputs to XSA stock assessments are correct. Second, we assume, consistently with some previous studies (Pascooe and Robinson, 1996; Millischer et al., 1999), that catchability is not constant over time, and has increased through fishing power development. Catchability trends are hereby described by an exponential model, similarly to the “Hybrid” method (Pope and Shepherd, 1985), but making the additional assumption that trends are kept equal across age ranges. Equation (1) may then be transformed into Equation (2)

\[ F(s,f,y,a) = q(s,f,a) e^{\alpha \ln(y - T + 1)}E(f,y) \]  

\[ F(s,f,y,a) = q(s,f,a) e^{\alpha \ln(y - T + 1)}E(f,y) \]  

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where $\alpha$, $T$ and $q$ respectively represent the annual growth in catchability, the final year of the tuning period and the catchability in the final year of the tuning period. Fishing effort is then adjusted by the exponential term of Equation (2), for the tuning fleets for which a trend in catchability has been detected by the single-fleet “Hybrid” assessment method.

The above methodology is applied to the assessment of cod, saithe, sole and plaice in the North Sea. The main tuning parameters, i.e. seasonal adjustment parameters, time series taper, ages below which catchability is dependent on stock size, ages above which catchability is independent of age, $F$ shrinkage, and minimal standard error of fleet estimates of population size (Shepherd, 1999) of the current assessment ($\alpha = 0\%$), are reported in ICES (2000). These tuning parameters are kept unchanged in the assessments performed with adjusted fishing effort, so as to highlight the focus on the catchability trends analysis.

**Performance criteria**

Performances of assessments operated with adjusted and non-adjusted fishing effort are compared using a number of criteria:

1. Improved precision of the estimates of Log-catchability for each selected tuning fleet
2. Reduced trends in the annual trajectories of the Log-catchability residuals for each selected tuning fleet
3. Enhanced consistency of the retrospective patterns

Criterion 1 is investigated by examining the coefficient of variation $CV$ relative to the Log-catchability estimates

$$CV(s,f,a) = \frac{\sigma(s,f,a)}{\ln[q(s,f,a)]}$$

where $\ln[q(s,f,a)]$ is the estimated value of Log-catchability and $\sigma(s,f,a)$ is the standard deviation associated with this estimate. Low CV should correspond to good agreement between the tuning fleet data and the populations as estimated by the XSA transformation of the catch data. $\sigma$ could have been used as an alternative criterion to $CV$. However, $\sigma$ is generally—as in the present assessments—not independent of the mean log-catchability.

Criterion 2 is investigated by examining the first-order auto-correlation $ACR$ of the Log-catchability residuals $\varepsilon(s,f,y,a)$

$$ACR(s,f,a) = \frac{COV(\varepsilon(s,f,y-1,a),\varepsilon(s,f,y,a))}{VAR(\varepsilon(s,f,y,a))}$$

where $COV$ and $VAR$ refer to the covariance and the variance functions respectively. Values of $ACR$ close to 1 would characterise trends in the residuals time series. Values of $ACR$ close to –1 would indicate that the residuals time series are oscillating around a stable mean, with a period of 2 years. Finally, values of $ACR$ close to 0 would identify a purely random process. Low trends in residuals are hence associated with values of $ACR$ included in the range [-1, 0].

Criterion 3 is investigated by examining the distance (retrospective index) between the annual trajectories relative to some variables of interest, derived from assessments performed in different years. The variables under examination are fishing mortality, $SSB$ and recruitment. The retrospective index associated with each of these variables is calculated through a double summation. The first summation sums, over the years of assessment $i = T_A, ..., T-1$, the relative distance between a variable in year $y$, as assessed in year $i$ $(i > y)$ and that variable in year $y$, as evaluated in the last assessment performed in year $T$. The scalar derived from the first summation is then summed over the data period $[T_0, T]$, resulting in a retrospective index. Should the different assessments be perfectly consistent, each retrospective index would equate 0. The retrospective indices $IF$, $IS$ and $IR$ respectively associated with these distances are formulated as

$$IF(s) = \sum_{y=T_0}^{T} \left[ \sum_{i=MAT(y,T_A)}^{T-1} \left( \frac{F(s,y,i) - F(s,y,T)}{F(s,y,T)} \right)^2 \right]$$

$$IS(s) = \sum_{y=T_0}^{T} \left[ \sum_{i=MAT(y,T_A)}^{T-1} \left( \frac{SSB(s,y,i) - SSB(s,y,T)}{SSB(s,y,T)} \right)^2 \right]$$

$$IR(s) = \sum_{y=T_0}^{T} \left[ \sum_{i=MAT(y,T_A)}^{T-1} \left( \frac{R(s,y,i) - R(s,y,T)}{R(s,y,T)} \right)^2 \right]$$

where $T_0$ and $T_A$ are respectively the first year of the retrospective period and the year of the first assessment, $F$ is the fishing mortality averaged over the reference fully recruited ages, and $R$ is the recruitment. Table 1 summarises information on the parameters used to calculate the three retrospective indices.
RESULTS

Table 2 shows the tuning fleets being prone to significant trends in catchability at age (slope higher than twice standard error), as derived from the “Hybrid” analysis. All the trends under consideration are positive and exceed 5% per year, which is a high level. The fact that no negative trends are identified for any age groups suggest that the positive trends observed are due to an overall increase in fishing power in some age groups, and not a change in the selectivity of the fleets. The trend in catchability of both English and Scottish seiners harvesting cod applies to the same age class. As a result, these tuning fleets have been aggregated for convenience into one tuning fleet, hereafter referred to as “UK
The high trend in catchability relative to the English beam-trawlers harvesting sole only applies to ages 11 and 12, which represent less than 5% of the catch in numbers. As a result, the English beam-trawlers harvesting sole are assumed to have constant catchability, for recruited ages, in subsequent XSA analyses. The only trend in catchability associated with saithe is detected for a survey. This unexpected outcome could be an artefact that would result from the English groundfish survey not being appropriate for estimating abundance indices relative to saithe. Overall, the tuning fleets harvesting saithe have constant catchability over time, and no subsequent analyses will be performed in relation to the saithe assessment.

Table 3 summarises information on the tuning fleets, for which a trend in catchability has been detected. The weight of the selected tuning fleets is
generally low (respectively high) in relation to young (respectively old) fish, and it is included in ranges 36-65% (cod), 9-67% (plaice) and 26-63% (sole). The remaining weight is taken by the other tuning fleets for which no catchability trend has been detected.

Figures 1a,b and 2a,b show that adjusting fishing effort with $\alpha$ enhances the precision in the Log-catchability estimates and reduces trends in the Log-catchability residuals for English otter-trawlers harvesting cod, but only for young age groups 2-4. The minimal CV and
ACR are both achieved at levels of $\alpha$ which roughly correspond to the values shown in Table 2. These results confirm that the fishing power of the English otter-trawl fishery has increased in relation to cod aged 2-4 years, which are the three most important age groups in terms of catch numbers at age (ICES, 2000). No dramatic improvements are brought about by adjusting the fishing effort of UK seiners harvesting cod (Figures 1c,d and 2c,d).

Figures 3a,b and 4a,b show that increasing $\alpha$ both increases precision of Log-catchability estimates and decreases trends in Log-catchability residuals for age groups 4-9 of plaice exploited by English beam-trawlers. This result bears out the conclusions drawn from Table 2. Figures 3c,d indicate that increasing $\alpha$ contributes little to improving precision of the Log-catchability estimates relative to Dutch beam-trawlers harvesting sole. Increasing $\alpha$ reduces trends of the residuals relative to age group 6 for the same fishery (Figs. 4c,d). This result is partially in agreement with the outcomes of Table 2.

Tables 4 and 5 show the retrospective indices relative to fishing mortality, SSB and recruitment of cod, plaice and sole, for the different assessments being examined. The lowest values of IF, IS and IR have been highlighted.

Figures 5, 6 and 7 contrast the retrospective patterns relative to fishing mortality, SSB and recruitment of cod, plaice and sole as derived from the standard assessment, with the retrospective patterns derived from the assessments providing the lowest IF, IS and IR (Tables 4 and 5).

Adjusting fishing effort overall does not affect the retrospective patterns in recruitment, for all stocks. IF and IS relative to cod are primarily affected by adjusting the fishing effort of the UK seiners. The assessment derived from combination (“20% x 8%”) provides the most consistent patterns in relation to SSB, but it does not really improve the retrospective pattern relative to fishing mortality. Adjusting the fishing effort of English beam-trawlers (respectively Dutch beam-trawlers) by applying a correction factor of $\alpha = 10\%$ enhances the overall assessment, by providing consistent retrospective patterns in relation to both fishing mortality and SSB of plaice (respectively sole).

Figure 8 represents the annual trajectories of fishing mortality (F), SSB and recruitment (R) derived from the North Sea cod, saithe, plaice and sole assessments, given as a function of catchability increase ($\alpha%$) for UK and Dutch beam-trawlers. The lowest values of IF, IS and IR are highlighted.

| Species | Tuning fleet | $\alpha\%$ | IF | IS | IR |
|-----------------|-----------------|-------|-----|-----|-----|
| Plaice | English Beam-trawlers | 0 | 0.17 | 0.13 | 0.73 |
| | | 2 | 0.12 | 0.10 | 0.73 |
| | | 4 | 0.09 | 0.07 | 0.73 |
| | | 6 | 0.07 | 0.06 | 0.73 |
| | | 8 | 0.05 | 0.05 | 0.73 |
| | 10 | 0.04 | 0.04 | 0.74 |
| Sole | Dutch Beam-trawlers | 0 | 0.05 | 0.04 | 0.15 |
| | | 2 | 0.03 | 0.03 | 0.14 |
| | | 4 | 0.02 | 0.02 | 0.13 |
| | | 6 | 0.02 | 0.01 | 0.12 |
| | | 8 | 0.01 | 0.01 | 0.11 |
| | 10 | 0.01 | 0.00 | 0.11 |

Table 4. – Retrospective indices relative to the cod assessment, given as a function of catchability increase ($\alpha\%$) for UK (English and Scottish) Seiners and English otter-trawlers. The lowest values of IF, IS and IR are highlighted.

| Species | Tuning fleet | $\alpha\%$ | IF | IS | IR |
|-----------------|-----------------|-------|-----|-----|-----|
| Plaice | English Beam-trawlers | 0 | 0.17 | 0.13 | 0.73 |
| | | 2 | 0.12 | 0.10 | 0.73 |
| | | 4 | 0.09 | 0.07 | 0.73 |
| | | 6 | 0.07 | 0.06 | 0.73 |
| | | 8 | 0.05 | 0.05 | 0.73 |
| | 10 | 0.04 | 0.04 | 0.74 |
| Sole | Dutch Beam-trawlers | 0 | 0.05 | 0.04 | 0.15 |
| | | 2 | 0.03 | 0.03 | 0.14 |
| | | 4 | 0.02 | 0.02 | 0.13 |
| | | 6 | 0.02 | 0.01 | 0.12 |
| | | 8 | 0.01 | 0.01 | 0.11 |
| | 10 | 0.01 | 0.00 | 0.11 |
sole assessments, performed with selected values of \( \alpha \). The \( \alpha \) values have been drawn from examination of Tables 4 and 5.

Ignoring annual increases in catchability overall results in an under-estimation of \( F \) and an over-estimation of both SSB and \( R \). Thus, while the current cod and plaice assessments diagnose an increase of SSB between 1997 and 1998, assessments performed with \( \alpha \) greater than 0% conclude that SSB has remained stable, at a different level (Fig. 8d). Likewise, while the current sole assessment diagnoses a reduction in \( F \) between 1997 and 1998, assessments performed with \( \alpha \) greater than 0% conclude that \( F \) has increased over the same period (Fig. 8c). Recruitment trajectories are not importantly modified by changes in \( \alpha \) (Figs. 8g-i).
The main conclusion of this study is that accounting for time variations in catchability would result in: (i) some improvement in the performances of traditional stock assessments; and (ii) new perceptions of the biomass and fishing mortality historical trends.

Adjusting the fishing effort of English beam-trawlers and Dutch beam-trawlers harvesting plaice and sole by the fishing power development estimated for these tuning fleets (about 10%) would increase the precision of catchability estimates, reduce trends in Log-catchability residuals and make the retrospective patterns relative to fishing mortality and spawning biomass more consistent.
By contrast with the standard assessment, assessments performed with adjusted fishing effort diagnose an increase in the fishing mortality of sole between 1997 and 1998.

Adjusting the fishing effort of English otter-trawlers harvesting cod by accounting for a 12% annual trend in fishing power would increase the precision of catchability estimates and reduce trends in Log-catchability residuals. Despite limited effects on the retrospective indices, the assessments performed with adjusted fishing effort diagnose, by contrast with the standard assessment, that the spawning biomass of cod remains stable between 1997 and 1998. Results also suggest that adjusting the fishing effort of the UK seiners harvesting would not improve the assessment performance much.

There are limitations to the present approach, and these are debated below.

Catchability dynamics have here been modelled through a simple single-parameterised exponential regression of fishing power over time. In fact, the time dynamics of catchability could be more complex. First, while fishing power could increase over given periods through technical development (Pascoe and Robinson, 1996), it would be expected to reach some kind of technological plateau in the future, and it could even decrease over other periods through management constraints (Gillis et al., 1995) or competition between fishing vessels (Gillis and Peterman, 1998). Second, it was assumed here that deterministic time dynamics in catchability only occurred through changes in fishing power. However, non-random variations may also occur through other components of catchability such as selectivity (Fryer, 1991; Sampson, 1993) and vulnerability (Arreguin-Sánchez and Pitcher, 1999). A more complex catchability model would give a better fit than the single-parameterised exponential model used in this study, but at the expense of a decreased robustness, and particularly a higher sensitivity to parameter uncertainty. It would appear difficult to measure the optimal level of complexity required for modelling changes in catchability over time. This task however, is outside the scope of this study.

Another limitation of the approach being undertaken in this study lies in the assumptions inherent to the other XSA inputs and parameters. Thus, we have assumed that the catch rates data were correct. In fact, the quality of these data could well be adversely altered as a result of mis-reporting, which is thought to occur, particularly in the case of cod and plaice (ICES, 2000). In addition, the reliability of catch rates as stock abundance index could be questioned as fish and fishing vessels are often subject to spatial aggregations (ICES, 2000). Finally, natural mortality is assumed to be constant over time. If, as evidence suggest, natural mortality was prone to non-random variation over time (Rice and Gislason, 1996), ignoring such dynamics would annihilate some of the benefits brought about by accounting for catchability dynamics, particularly with regard to fishing mortality estimates.

Further research could be carried out in several directions. First, a more comprehensive model of catchability dynamics could be developed so as to standardise the fishing effort of the tuning fleets more accurately. The catchability model, possibly combined with another model describing some of the dynamics in natural mortality (e.g. ICES, 1997), could then, if successfully tested against existing data, be considered as an auxiliary module to be included in ICES assessment packages. Second, the present approach could be further expanded to examine how sensitive the outcomes of medium-term projections and the definition of biological reference points are to uncertainty in catchability dynamics.

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