Mental Health During the First Weeks of the COVID-19 Pandemic in the United States

William D. S. Killgore*, Sara A. Cloonan, Emily C. Taylor and Natalie S. Dailey

Social, Cognitive, and Affective Neuroscience Lab, Department of Psychiatry, University of Arizona College of Medicine, Tucson, AZ, United States

Background: By March 2020, the World Health Organization declared the COVID-19 crisis as a worldwide pandemic and many local governments instituted stay-at-home orders and closed non-essential businesses. Within the United States, tens of millions of workers lost their jobs and financial security during the first few weeks of the national response, in an attempt to slow the global pandemic. Because of the enormity of the pandemic and its potential impact on mental health, the objective of the present study was to document the prevalence of mental health problems and their association with pandemic-related job loss during the third week of the nationwide shutdown.

Methods: Mental health was assessed via online questionnaires among a representative sample of 1,013 U.S. adults on April 9–10, 2020. Rates of clinically significant mental health outcomes were compared between participants who lost their job as a result of COVID-19 restrictions (17.4%) vs. those who did not (82.6%). Bivariate multiple logistic regression identified factors that were predictive of, and protective against, mental health problems.

Results: The prevalence of clinically significant symptoms was significantly higher than prior population estimates, ranging from 27 to 32% for depression, 30 to 46% for anxiety disorders, 15 to 18% for acute/post-traumatic stress, 25% for insomnia, and 18% for suicidal ideation. Prevalence estimates were 1.5–1.7 times higher for those who reported job loss due to COVID-19 restrictions than those who did not. Mental health problems were predicted by worry over financial instability, insomnia, social isolation, and alcohol consumption, while getting outside more often, perceived social support, and older age were protective against these problems.

Conclusions: During the first 3 weeks of lockdowns/stay-at-home restrictions, mental health problems, including depression, anxiety, insomnia, and acute stress reactions were notably elevated relative to prior population estimates. Job loss related to the nationwide shutdown was particularly associated with poorer mental health. These findings provide a baseline of mental health functioning during the first weeks of the national emergency and lockdown orders in response to COVID-19.

Keywords: COVID-19, depression, generalized anxiety, PTSD, insomnia, mental health, job loss, financial worries
INTRODUCTION

During the first weeks of the COVID-19 pandemic, nations around the globe implemented unprecedented measures to mitigate the transmission of the SARS CoV-2 virus (1, 2). The United States reported the first COVID-19-related death on February 29, 2020, and <2 weeks later, a national state of emergency was called. In response, local and state governments closed schools, shut down all non-essential business, and enacted shelter-in-place orders, with the first state-wide shutdown occurring in California on March 19, 2020. Within a matter of weeks, every state in the country had enacted some form of restriction (3), with most of the U.S. population asked to remain at home and to severely limit physical proximity to others. As a consequence, most employment activities deemed “non-essential” by local governments had ceased or rapidly shifted to remote telecommuting or work-from-home options. This soon led to large scale furloughs and job losses for a large segment of the population (4). Despite the clear public health necessity of the stay-at-home orders and physical distancing strategies to slow the spread of the virus, there is no question that these efforts profoundly altered the basic foundations of the social and occupational lives of much of the population. Consequently, the potential long-term effects of the pandemic and associated restrictions on mental health will likely be a focus of research for years to come.

The financial consequences of the shelter-in-place mitigation strategies were rapidly felt throughout the country. Within the first 6 weeks of state-wide stay-at-home orders in the U.S., more than 33 million Americans had filed new unemployment claims, a level of job loss that had not been seen since the Great Depression (4). This extraordinary surge in unemployment was troubling in light of the well-established findings that job loss, financial stresses, and lack of social support are leading contributors to suicide, substance abuse, domestic violence, and other mental health issues (5, 6). The rampant uncertainty surrounding the potential course of the pandemic and widespread concerns over financial instability stemming from the lockdown orders led to rising concerns that a surge in mental health problems may be looming on the horizon (7, 8). Fear of the virus, its transmissibility, and its potential lethality contributed to panic and generalized anxiety (9), and raised concerns that post-traumatic stress symptoms could persist long after the pandemic had resolved (10), as has been seen in other countries (11). Quarantines enacted in prior disease outbreaks have also been shown to significantly elevate symptoms of PTSD and depression among the population (12). Moreover, prolonged stay-at-home requirements and social distancing measures may have unintended mental health consequences, as they restrict many of the facets of daily existence that provide emotional resilience, social connection, and satisfaction with life (13). The enormity of the pandemic and its effects on daily existence led many experts to voice concern that mental health problems may be a lingering issue for years to come (7, 14).

Effective recovery from the pandemic will require a well-documented and comprehensive understanding of the mental health effects that emerged during the acute stages of the crisis. To that end, we studied the point prevalence of mental health problems in a nationally representative sample of adults in the U.S. collected at the very outset of the pandemic, during the third week of the nationwide stay-at-home restrictions. We identified overall rates of mental health concerns in this sample and determined the differences in mental health outcomes between those who had lost their primary employment due to the economic shutdown and those who had not. We believe that these data will be critical for documenting the mental health status of the population during the initial phase of the COVID-19 pandemic and will serve as a benchmark for future research on the long-term psychiatric outcomes of the crisis.

METHODS

Participants

Here, we summarize mental health data from an online assessment collected over a 28-h period between April 9 and April 10, 2020. This date was selected because it was, at that time, projected to be the peak of expected U.S. deaths due to COVID-19 according to the University of Washington Institute for Health Metrics and Evaluation (IHME; https://covid19.healthdata.org/united-states-of-america) model during the preceding week. Using the Amazon Mechanical Turk (MTurk) online crowdsourcing platform (15), a total of 1,074 participants were screened and provided small financial compensation for their time. All participants were geographically located within the United States (verified by IP address geo-coordinates), were at least 18 years of age, and reported English as their primary language. A brief screen for reading comprehension excluded 48 volunteers from further participation. The remaining 1,026 individuals then completed the online questionnaires. Data from 13 participants were excluded for failing to correctly answer imbedded attention check questions. This resulted in a final sample of 1,013 participants with complete and valid data, which formed the basis of the present analysis. The sampling of participants from each state was closely proportional to state population according to the 2019 U.S. Census, suggesting a nationally representative sample of U.S. adults. Specifically, we calculated the proportion of participants from each state, relative to the national population total, for each state and for the U.S. Census data for each state. The mean absolute difference in sampling proportions and census data proportions across states was 0.004 (i.e., <0.5% point). Further, we found that all sampled state proportions differed from the census data by 1.5% points or less, except for Texas (which was underrepresented in the current sample by 3% points). The intraclass correlation coefficient (ICC) between the sample proportions and the census data across states was quite high (ICC = 0.95, p < 0.0001), suggesting that these data are likely representative of the larger population. All participants provided electronic acknowledgment of informed consent after begin provided with a full description of the study. This study was approved by the institutional review board of the University of Arizona.
Questionnaires and Primary Outcomes

Our goal was to provide a source of reliable documentation of the initial mental health prevalence estimates during the first weeks of the pandemic response, which would provide a benchmark for future research efforts. In an effort to be comprehensive, we included several outcome metrics that assessed similar constructs (e.g., we collected two measures of depressive mood, one longer and more extensive, and another as a brief screener). Outcomes were focused on major mental health symptoms that could potentially result from concerns surrounding the COVID-19 pandemic and/or societal attempts to mitigate the spread of the illness. These included current symptoms (i.e., present within the last week to month) of a major depressive disorder, generalized anxiety disorder (GAD), and acute/post-traumatic stress disorder. Using established instruments, we calculated mean scores and the percentage of the sample exceeding published cut-off scores. Depressive symptoms were assessed with the Beck Depression Inventory-II (≥20) (16) and the Patient Health Questionnaire-9 (≥10) (17). We also examined separately the suicidal ideation scores on each of these depression scales (i.e., Item 9 scores ≥1). Generalized anxiety disorder (GAD) symptoms were measured with the GAD-7 (≥8) (18, 19), the Zung Self-Rated Anxiety Scale (SRAS; ≥36) (20, 21), and the state and trait portions of the Spielberger State-Trait Anxiety Inventory (STAI; ≥46) (22, 23). To assess acute/post-traumatic stress, we administered the National Stressful Events Survey Acute Stress Disorder Short Scale (NSESSS; ≥2) (24, 25), and the Primary Care Post-Traumatic Stress Disorder (PC-PTSD) scale (≥3) (26, 27). We also measured the severity of insomnia symptoms using the Insomnia Severity Index (ISI; ≥15) (28, 29). To examine potential moderators of the mental health outcomes described above, we also collected data on a variety of demographic factors, particularly related to job loss from the shutdown and socioeconomic status, and specific concerns relevant to the ongoing COVID-19 pandemic. These variables are detailed in Table 1.

Statistical Analysis

We proposed an initial sample size of 1,000 participants, which would provide 88% power to detect effects of interest at a two-sided significance criterion of α = 0.05, assuming small effect sizes. The data were analyzed using SPSS software (version 26). Means and descriptive statistics were calculated for the sample as a whole, as well as for subgroups of individuals who reported job loss as a result of the societal responses aimed at mitigating the spread of COVID-19. Mean values between job status groups were compared with analysis of covariance (ANCOVA) statistically controlling for pre-pandemic income level, education level, and potential exposure to COVID-19 (i.e., “have you noticed that you are shown symptoms of COVID-19?” and “has anyone in your household been diagnosed with COVID-19?”). For clinically relevant scales where published cut-off points were available, we calculated the percentage of the sample and subgroups that met or exceeded those values. Chi-squared statistics were calculated to compare the percentage of individuals exceeding the cut-offs in each group. Finally, binomial multiple logistic regression analyses, with forward selection using the Likelihood Ratio, were used to identify key concerns, traits, and behavioral factors associated with meeting criteria for a probable mental health issue. Multiple comparisons were controlled by false discovery rate (FDR) adjustment of significance, which was calculated using the Benjamini-Hochberg procedure, as implemented via the online FDR (tool: https://tools.carbocation.com/FDR). The corrections were applied uniformly for all comparisons within each table.

RESULTS

Demographics

Primary demographic characteristics of the sample were well-matched to the larger U.S. population (see Methods section). While an upper restriction on age was not set, those who responded to the survey included adults ranging in age from 18 to 82 years and females were slightly over-represented (i.e., 56.4%). Overall, 65.9% of the sample reported a previous year household income of $75,000 or less.

Overall Mental Health Outcomes

For the sample as a whole, the proportions of individuals screening positive for probable mental health problems was notably higher than would be expected in a similar sample based on prior research (30, 31). For those screening positive for moderate to severe depression (Figure 1, Table 2), the prevalence ranged from 27.2% (BDI-II) to 32.2% (PHQ-9). Across four different scales commonly used to screen for GAD and other clinically significant anxiety disorders, the screen-positive prevalence ranged from 29.8% (STAI-State) to 45.8% (SRAS). Further, 15.2% of the sample screened positive for probable Acute Stress Disorder (NSESSS), while 17.9% screened positive on a brief measure of possible PTSD symptoms (PC-PTSD). Critically, 17.6% of the sample screened positive for some evidence of suicidal ideation on two different scales (BDI-II Item 9; PHQ-9 Item 9), while 25.1% of the sample screened positive for clinically significant insomnia on the ISI.

COVID-19 Job Loss and Mental Health Outcomes

When asked the question “Have you lost your primary job/income due to COVID-19?”, 17.4% of the respondents answered “yes,” while 82.6% answered “no.” As evident in Table 2, individuals who lost their job due to COVID-19 scored significantly higher on measures of depression (BDI-II and PHQ-9). Moreover, 38.1% of those who reported job loss due to COVID-19 exceeded the cut-off for moderate to severe depression on the BDI-II, whereas 25.2% of those who had not lost their job met that criterion. This was even more notable for the PHQ-9, with 44.3% of those who lost their job scoring in the clinically significant range, compared to 29.6% who did not. These differences were significant even after controlling for pre-pandemic income level, education, and the perception of close exposure to COVID-19.

Similar patterns were observed for measures of GAD and other anxiety disorders. Most notably, 57.4% of those who lost their job during the pandemic met the cut-off for a probable
### TABLE 1 Demographic characteristics and COVID-19 concerns of groups of individuals who lost or did not lose their primary employment due to the pandemic.

| Characteristics                                      | Total sample (N = 1,013) | COVID-19 job loss (N = 176) | No job loss (N = 837) | p-Value |
|-----------------------------------------------------|--------------------------|-----------------------------|-----------------------|---------|
| *Age—yr                                             | 36.74 ± 12.09            | 34.80 ± 10.43               | 37.14 ± 12.38         | 0.009   |
| *Female sex—no. (%)                                 | 567 (56.4)               | 110 (62.9)                  | 457 (55.0)            | n.s.    |
| *Education—yr                                       | 15.0 ± 2.1               | 14.6 ± 2.0                  | 15.1 ± 2.1            | 0.004†  |
| *Ethnicity—no. (%)                                  |                          |                             |                       |         |
| White                                               | 776 (76.8)               | 139 (79.0)                  | 637 (76.1)            |         |
| Black/African American                               | 99 (9.8)                 | 19 (10.8)                   | 80 (9.6)              |         |
| Hispanic/Latino                                      | 43 (4.2)                 | 5 (2.8)                     | 38 (4.5)              |         |
| Asian                                                | 66 (6.5)                 | 8 (4.5)                     | 58 (6.9)              |         |
| Native American/American Indian/Alaska Native        | 5 (0.5)                  | 1 (0.6)                     | 4 (0.5)               |         |
| Native Hawaiian/Pacific Islander                     | 1 (0.1)                  | 0 (0.0)                     | 1 (0.1)               |         |
| Other                                                | 21 (2.1)                 | 4 (2.3)                     | 17 (2.0)              |         |
| Prefer not to answer                                 | 2 (0.2)                  | 0 (0.0)                     | 2 (0.2)               |         |
| *Income—no. (%)                                      |                          |                             |                       | 0.013   |
| ≤$10,000                                            | 50 (4.9)                 | 10 (5.7)                    | 40 (4.8)              |         |
| $10,001–$25,000                                     | 112 (11.1)               | 30 (17.1)                   | 82 (9.8)              |         |
| $25,001–$50,000                                     | 267 (26.4)               | 53 (30.3)                   | 214 (25.6)            |         |
| $50,001–$75,000                                     | 238 (23.5)               | 35 (20.0)                   | 203 (24.3)            |         |
| $75,001–$100,000                                     | 172 (17.0)               | 31 (17.7)                   | 141 (16.8)            |         |
| $100,001–$150,000                                    | 122 (12.1)               | 11 (6.3)                    | 111 (13.3)            |         |
| $150,001–$200,000                                    | 34 (3.4)                 | 4 (2.3)                     | 30 (3.6)              |         |
| ≥$200,001                                           | 17 (1.7)                 | 1 (0.6)                     | 16 (1.9)              |         |
| *COVID-19 related concerns—no. reporting YES (%)    |                          |                             |                       |         |
| Have you noticed that you are showing symptoms of COVID-19 (fever, dry cough, fatigue/soreness)? | 113 (11.2)               | 29 (16.5)                   | 84 (10.0)             | 0.014†  |
| Have you been tested for COVID-19?                   | 31 (3.1)                 | 8 (4.5)                     | 23 (2.7)              | n.s.    |
| Have you been formally diagnosed with COVID-19?      | 4 (0.4)                  | 2 (1.1)                     | 2 (0.2)               | n.s.    |
| Are you considered to be in a “high-risk” group for COVID-19? | 309 (30.5)               | 57 (32.4)                   | 252 (30.1)            | n.s.    |
| Has anyone in your household (i.e., where you live) been diagnosed with COVID-19? | 16 (1.6)                 | 7 (4.0)                     | 9 (1.1)               | 0.005†  |
| Have any of your friends, co-workers, or first-degree relatives been diagnosed with COVID-19? | 203 (20.0)               | 47 (26.7)                   | 156 (18.6)            | 0.015†  |
| Do you know anyone personally who has been diagnosed with COVID-19? | 344 (34.0)               | 75 (42.6)                   | 269 (32.1)            | 0.008†  |
| Do you know anyone personally who has died from complications associated with COVID-19? | 74 (7.3)                 | 17 (9.7)                    | 57 (6.8)              | n.s.    |
| Are you currently “sheltering in place” (i.e., not leaving home except for necessities)? | 948 (93.6)               | 171 (97.2)                  | 777 (92.8)            | 0.033   |
| Have you become worried about your ability to financially support yourself and loved ones? | 555 (54.8)               | 157 (89.2)                  | 398 (47.6)            | <0.001† |
| Do you have someone you care about or who is emotionally close to you that you can talk to daily? | 921 (90.9)               | 162 (92.0)                  | 759 (90.7)            | n.s.    |
| Do you feel socially isolated?                       | 579 (57.2)               | 122 (69.3)                  | 457 (54.8)            | <0.001† |
| Do you feel like you have enough social/emotional support to get through this time? | 844 (83.3)               | 128 (72.7)                  | 716 (85.5)            | <0.001† |
| Are you engaging in consistent “social distancing” around people (e.g., keeping 6 feet from others)? | 977 (96.4)               | 174 (98.9)                  | 803 (96.9)            | n.s.    |
| Are you avoiding all contact with others outside of the home? | 865 (85.4)               | 155 (88.1)                  | 710 (84.8)            | n.s.    |
| Do you touch others less?                           | 957 (94.5)               | 167 (94.9)                  | 790 (94.4)            | n.s.    |
| Do you trust other people less?                     | 519 (51.2)               | 102 (58.0)                  | 417 (49.8)            | 0.05    |
| How often do you pray?                              |                          |                             |                       | n.s.    |

(Continued)
TABLE 1 | Continued

| Characteristics | Total sample (N = 1,013) | COVID-19 job loss (N = 176) | No job loss (N = 837) | p-Value |
|-----------------|--------------------------|-----------------------------|----------------------|---------|
| At least once a day | 270 (27.4) | 49 (29.0) | 221 (27.0) |         |
| 1–6 days a week | 113 (11.4) | 19 (11.2) | 94 (11.5) |         |
| At least monthly | 119 (12.1) | 24 (14.2) | 95 (11.6) |         |
| Seldom or Never | 485 (49.1) | 77 (45.6) | 408 (49.9) |         |
| *Weekly exercise—min | 36.2 ± 43.2 | 36.1 ± 46.3 | 36.2 ± 42.5 | n.s. |
| *Over the past two weeks, how many days did you get outside during sunlight hours for more than 10 minutes? | 4.7 ± 2.1 | 4.8 ± 2.0 | 4.7 ± 2.1 | n.s. |
| *On average, how many minutes did you spend outside in the sun light each day? | 66.0 ± 68.0 | 68.0 ± 64.3 | 65.5 ± 68.8 | n.s. |
| *How often do you have one drink containing alcohol?—% ≥2 times per wk | 25.6% | 21.0% | 26.5% | n.s. |
| *How many drinks containing alcohol do you have on a typical day when you are drinking?—% ≥3 per session | 30.2% | 35.7% | 29.1% | n.s. |

1 Significant at false discovery rate (FDR) correction, p < 0.05; n.s., non-significant; yr, year; no., number.

*Items included as potential predictors of mental health in the binomial logistic regression (as well as job loss category and insomnia severity, which are not listed).

anxiety disorder on the Zung SRAS, compared to 43.3% who had not lost their job.

Evidence of probable acute stress and post-traumatic stress reactions was evident in 22.9% and 25.0% of those who lost a job due to the coronavirus outbreak, respectively, compared to 13.5% and 16.4%, respectively, among those whose jobs were not affected.

While suicidal ideation on Item 9 of the BDI-II was not different between job-loss groups, there was a significant difference in suicidal ideation on Item 9 of the PHQ-9, with 25% of those who reported losing their job endorsing some level of suicidal ideation, compared to 16% of those who had not.

Finally, the data in Table 2 also show that 32.4% of those who lost their job met or exceeded the cut-off for clinically significant insomnia on the ISI, while 23.5% of those who had not lost their job met this level of severity.

Contributing Factors to Mental Health Outcomes

To identify some potential factors that may mitigate or exacerbate mental health problems during the pandemic, we queried participants on a series of questions related to concerns about COVID-19 (see Table 1). While it is acknowledged that no single set of variables will provide a comprehensive explanation of mental health issues, we selected a set of items focused on COVID-19 concerns that we believed would likely play a role in mental health responses to the pandemic. This approach, using background knowledge to aid in variable selection is considered to be an accepted approach to regression analysis (32). For the sample as a whole, a total of 28 potential variables of interest were initially included based on their relevance to pandemic-related concerns at the time, including fears of the virus itself, being in close proximity to someone with the virus, perceived social support, daily activities, alcohol consumption, and basic demographics. The analysis included the asterisked items in Table 1, as well as variables assessing job loss due to the pandemic and insomnia severity. These 28 variables were entered into a binomial multiple logistic regression with forward selection to predict the likelihood of meeting positive screening criteria for each of the 10 mental health outcome variables with published cut-offs (see Table 2). The variables surviving selection for each model are listed in Table 3. The most significant predictors of meeting criteria for moderate to severe depression on the BDI-II and the PHQ-9 included worry about the ability to financially support oneself or family, feeling socially isolated, and greater alcohol use, while spending more days each week outside in the sunshine and feeling that one had enough social support appeared protective against depression. As evident in Table 3, while there were a number of factors associated with screening positive for an anxiety disorder on the GAD-7, Zung SRAS, STAI-S, and STAI-T, the most consistent predictors across measures included endorsing worry about the ability to financially support self or family, feeling socially isolated, and problems with insomnia, while spending more days per week outside in the sunshine appeared most consistently protective. The probability of meeting criteria for acute stress or post-traumatic stress reaction on the NSESSS or PC-PTSD was greatest among those reporting worry about financial problems, feeling socially isolated, trusting others less, and endorsing more problems with insomnia. Finally, screening positive for suicidal ideation on Item 9 of the BDI-II and PHQ-9 was most associated with endorsing financial worries, greater alcohol use, and problems with insomnia, while protective factors against suicidal ideation included male sex, older age, and feeling that one had enough social support to get through the crisis.

DISCUSSION

The reported prevalence of mental health problems during the first weeks of the pandemic response in the U.S. was notably higher than expected based on general population
FIGURE 1 | Histograms showing the distribution of scores on the eight major mental health assessment questionnaires. (A) Beck Depression Inventory-II (BDI); (B) Patient Health Questionnaire-9 (PHQ 9); (C) Generalized Anxiety Disorder-7 (GAD 7); (D) Zung Self-Rated Anxiety Scale; (E) Spielberger State Trait Anxiety Inventory-State (STAI State); (F) Spielberger State Trait Anxiety Inventory-Trait (STAI Trait); (G) NSESSS Acute Stress Disorder Short Scale; (H) Primary Care PTSD Screen for DSM-5 (PC-PTSD). Bars represent the percentage of participants obtaining a particular score. Each histogram divides the sample into those who met published cut-off points for clinical significance (blue, normal range; red, clinical range).
TABLE 2 | Mean scores and percent meeting clinical cut-offs on mental health outcomes between groups of individuals who lost or did not lose their primary employment due to the pandemic.

| Outcome | Total sample (N = 1,013) | COVID-19 job loss (N = 176) | No job loss (N = 837) | p-Value |
|---------|--------------------------|-----------------------------|----------------------|---------|
| BDI     | Clinically significant (BDI ≥ 20) ([Beck, 1996 #3123]) | 13.6 ± 12.0 | 17.5 ± 12.8 | 12.8 ± 11.7 | <0.001† |
|         | Mean total | 278 (27.4) | 67 (38.1) | 211 (25.2) | 0.001† |
| PHQ9    | Clinically significant (PHQ9 ≥ 10) ([Kroenke, 2001 #5911]) | 7.1 ± 6.3 | 9.2 ± 6.9 | 6.6 (6.1) | <0.001† |
|         | Mean total | 362 (32.2) | 78 (44.3) | 589 (29.6) | <0.001† |
| GAD7    | Clinically significant (GAD7 ≥ 6) | 6.0 ± 5.8 | 8.0 ± 6.2 | 5.6 ± 5.6 | <0.001† |
| Zung SRAS | Clinically significant (SRAS ≥ 8) ([Plummer, 2016 #5914]) | 321 (32.7) | 79 (45.4) | 242 (29.9) | <0.001† |
| STAI-state | Clinically significant (STAI-S ≥ 46) ([Fisher, 1999 #5916]) | 403 ± 11.0 | 44.8 ± 11.3 | 39.3 ± 10.8 | <0.001† |
| STAI-trait | Mean total | 302 (29.8) | 80 (45.5) | 222 (26.6) | <0.001† |
| NSESSS  | Clinically significant (SASS ≥ 2) ([Kilpatrick, 2013 #5917]) | 6.1 ± 6.3 | 7.9 ± 6.8 | 5.7 ± 6.1 | 0.001† |
| PC-PTSD | Clinically significant (PC-PTSD ≥ 3) ([Ouimette, 2008 #5910]) | 1.0 ± 1.4 | 1.3 ± 1.5 | 0.9 ± 1.3 | 0.014† |
| Suicidal ideation (BDI Item 9) | Mean total | 0.2 ± 0.7 | 0.2 ± 0.5 | 0.2 ± 0.5 | n.s. |
|         | Clinically significant (BDI Item 9 ≥ 1) | 178 (17.6) | 37 (21.0) | 141 (16.8) | n.s. |
| Suicidal ideation (PHQ9 Item 9) | Mean total | 0.3 ± 0.5 | 0.4 ± 0.8 | 0.3 ± 0.7 | n.s. |
|         | Clinically significant (PHQ Item 9 ≥ 1) | 178 (17.6) | 44 (25.0) | 134 (16.0) | 0.004† |
| ISI     | Mean total | 9.5 ± 6.8 | 11.1 ± 6.9 | 9.2 ± 6.7 | 0.011† |
|         | Clinically significant (SI ≥ 15) | 254 (25.1) | 57 (32.4) | 197 (23.4) | 0.014† |

n.s., non-significant; yr, year; no., number; BDI, Beck Depression Inventory; PHQ, Patient Health Questionnaire; GAD7, Generalized Anxiety Disorder; SRAS, Self-Reported Anxiety Scale; STAI, State-Trait Anxiety Inventory; NSESSS, National Stressful Events Survey Acute Stress Disorder Short Scale; PC-PTSD, Primary Care Post-traumatic Stress Disorders inventory; ISI, Insomnia Severity Index. †Significant at false discovery rate (FDR) correction, p < 0.05; Mean comparisons corrected for prior income, formal educational attainment, and potential exposure to COVID-19 (i.e., personally showing symptoms of COVID-19; someone in household diagnosed with COVID-19).

estimates collected over the decade prior to the pandemic. Prior research has estimated that the 12-month prevalence for mental health problems in the general population to be approximately 9.3% for any major depressive episode, 2.9% for GAD, and 4.4% for PTSD (30), while reported suicidal ideation ranges from 2.0 to 3.7% (31). Our findings suggest that the prevalence of probable mental health problems at this early phase of the pandemic was higher than estimates from prior years. Consistent with other contemporaneous research (33), we found that major depression was 2.9–3.5 times higher than before the pandemic; GAD 10.3–15.8 times higher; and acute-stress/post-traumatic stress 3.5–4.0 times higher. Moreover, suicidal ideation was found to be 4.8–8.8 times higher than prior population estimates. It should be kept in mind that brief screening methods such as those used here may tend to overestimate mental health prevalence rates relative to gold-standard clinical interviews (34), and that the comparison data were, in many cases, collected years earlier. Nonetheless, with due consideration to this risk, the magnitude of the findings raise serious concerns about the mental health status of the general U.S. population during the early phase of the COVID-19 pandemic.

As a result of necessary and vital efforts to slow the spread of the novel coronavirus, non-essential businesses were closed
TABLE 3 | Factors contributing to total sample mental health outcomes based on binomial logistic regression.

|                          | Nagelkerke R² | β (SE) | p-Value   | Adjusted odds ratio (95% CI) |
|--------------------------|---------------|--------|-----------|-----------------------------|
| BDI depression           | 0.39†         |        | <0.001    | 2.45 (1.70–3.54)            |
| Worry about ability to financially support self/family | 0.90 (0.19)  | <0.001 | 2.45 (1.70–3.54) |
| Feel socially isolated   | 0.50 (0.19)  | 0.011  | 1.64 (1.12–2.40) |
| Alcoholic drinks per day | 0.25 (0.10)  | 0.013  | 1.28 (1.05–1.55) |
| Insomnia severity index score | 0.13 (0.01) | <0.001 | 1.13 (1.10–1.17) |
| Age—yr                  | -0.03 (0.01) | <0.001 | 0.97 (0.96–0.99) |
| Days outside in sunlight per week—no. | -0.10 (0.04) | 0.012  | 0.90 (0.83–0.98) |
| Have enough social support to get through this | -1.07 (0.21) | <0.001 | 0.34 (0.22–0.52) |
| PHQ-9 depression         | 0.46†         |        | <0.001    | 1.83 (1.28–2.61)            |
| Worry about ability to financially support self/family | 0.60 (0.18)  | 0.001  | 1.78 (1.23–2.60) |
| Feel socially isolated   | 0.58 (0.19)  | 0.003  | 1.42 (1.17–1.73) |
| Alcoholic drinks per day | 0.35 (0.10)  | <0.001 | 1.20 (1.17–1.24) |
| Days outside in sunlight per week—no. | -0.18 (0.04) | <0.001 | 0.84 (0.77–0.91) |
| Touch others less        | -0.78 (0.36) | 0.029  | 0.46 (0.23–0.92) |
| Have enough social support to get through this | -0.81 (0.23) | <0.001 | 0.45 (0.29–0.69) |
| GAD-7 generalized anxiety disorder | 0.42†         |        | <0.001    | 2.33 (1.64–3.30)            |
| Worry about ability to financially support self/family | 0.84 (0.18)  | <0.001 | 2.13 (1.49–3.05) |
| Feel socially isolated   | 0.76 (0.18)  | <0.001 | 1.75 (1.25–2.47) |
| Female sex               | 0.56 (0.17)  | 0.001  | 1.51 (1.08–2.12) |
| Trust others less        | 0.41 (0.17)  | 0.017  | 1.17 (1.14–1.21) |
| Insomnia severity index score | 0.16 (0.01) | <0.001 | 0.97 (0.96–0.99) |
| Days outside in sunlight per week—no. | -0.11 (0.04) | 0.005  | 0.89 (0.82–0.97) |
| Anyone in household been diagnosed with COVID-19 | -1.54 (0.65) | 0.019  | 0.21 (0.06–0.78) |
| Zung SRAS anxiety disorder | 0.47†         |        | <0.001    | 2.24 (1.62–3.09)            |
| Worry about ability to financially support self/family | 0.81 (0.16)  | <0.001 | 2.02 (1.46–2.78) |
| Trust others less        | 0.70 (0.16)  | <0.001 | 1.65 (1.20–2.29) |
| Female sex               | 0.50 (0.17)  | 0.026  | 1.46 (1.05–2.03) |
| Know someone personally diagnosed with COVID-19 | 0.38 (0.17)  | 0.026  | 1.44 (1.03–2.00) |
| Feel socially isolated   | 0.36 (0.17)  | 0.031  | 1.25 (1.03–1.52) |
| Insomnia severity index score | 0.22 (0.10) | 0.026  | 1.19 (1.16–1.23) |
| Days outside in sunlight per week—no. | -0.15 (0.04) | <0.001 | 0.86 (0.80–0.93) |
| Touch others less        | -0.98 (0.35) | 0.005  | 0.38 (0.19–0.74) |
| STAI-state anxiety disorder | 0.36†         |        | <0.001    | 2.87 (2.03–4.08)            |
| Worry about ability to financially support self/family | 1.06 (0.18)  | <0.001 | 2.07 (1.44–2.98) |
| Feel socially isolated   | 0.73 (0.19)  | <0.001 | 1.61 (1.15–2.24) |
| Female sex               | 0.48 (0.17)  | 0.005  | 1.14 (1.11–1.17) |
| Know someone personally diagnosed with COVID-19 | 0.13 (0.01)  | <0.001 | 0.90 (0.84–0.97) |
| Days outside in sunlight per week—no. | -0.10 (0.04) | 0.008  | 0.68 (0.48–0.97) |
| Insomnia severity index score | -0.39 (0.18) | 0.035  | 0.64 (0.43–0.97) |
| Have enough social support to get through this | -0.44 (0.21) | 0.037  | 0.38 (0.19–0.74) |
| STAI-trait anxiety disorder | 0.35†         |        | <0.001    | 1.97 (1.44–2.68)            |
| Worry about ability to financially support self/family | 0.68 (0.16)  | <0.001 | 1.96 (1.42–2.71) |
| Feel socially isolated   | 0.67 (0.17)  | <0.001 | 1.25 (1.11–1.40) |
| Prayer frequency         | 0.22 (0.06)  | <0.001 | 1.11 (1.08–1.14) |
| Min of exercise per week | -0.01 (0.00) | 0.008  | 0.99 (0.99–1.00) |

(Continued)
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TABLE 3 | Continued

| Nagelkerke $R^2$ | $\beta$ (SE) | p-Value | Adjusted odds ratio (95% CI) |
|------------------|--------------|---------|----------------------------|
| Age—yr           | −0.02 (0.01) | 0.009   | 0.98 (0.97–1.00)           |
| Days outside in sunlight per week—no. | −0.14 (0.04) | <0.001 | 0.87 (0.81–0.94)           |
| Touch others less | −0.71 (0.32) | 0.026   | 0.49 (0.26–0.92)           |
| Have enough social support to get through this | −0.75 (0.22) | 0.001   | 0.47 (0.31–0.72)           |
| NSESSS acute stress disorder | 0.34† | <0.001 |                       |
| Been tested for COVID-19 | 1.15 (0.45) | 0.010   | 3.17 (1.32–7.59)           |
| Trust others less | 0.80 (0.23) | <0.001 | 2.22 (1.42–3.49)           |
| Worry about ability to financially support self/family | 0.75 (0.24) | 0.002   | 2.11 (1.32–3.39)           |
| Feel socially isolated | 0.69 (0.25) | 0.005   | 2.00 (1.23–3.27)           |
| Insomnia severity index score | 0.14 (0.02) | <0.001 | 1.15 (1.11–1.19)           |
| Age—yr           | −0.05 (0.01) | <0.001 | 0.96 (0.94–0.98)           |
| Prayer frequency  | 0.18 (0.08) | 0.030   | 2.11 (0.71–0.98)           |
| Currently sheltering in place | −0.79 (0.38) | 0.038   | 0.45 (0.21–0.96)           |
| PC-PTSD screen positive | 0.28† | <0.001 |                       |
| Feel socially isolated | 0.74 (0.22) | 0.001   | 2.10 (1.36–3.23)           |
| Worry about ability to financially support self/family | 0.70 (0.21) | 0.001   | 2.01 (1.33–3.06)           |
| Showing symptoms of COVID-19 | 0.69 (0.25) | 0.005   | 1.99 (1.23–3.23)           |
| Know someone personally who died from COVID-19 | 0.61 (0.31) | 0.049   | 1.84 (1.00–3.37)           |
| Trust others less | 0.51 (0.20) | 0.011   | 1.66 (1.12–2.46)           |
| Insomnia severity index score | 0.11 (0.01) | <0.001 | 1.12 (1.09–1.15)           |
| BDI suicidal ideation | 0.23† | <0.001 |                       |
| Worry about ability to financially support self/family | 0.64 (0.20) | 0.002   | 1.89 (1.27–2.81)           |
| Alcohol drinks per day | 0.33 (0.10) | 0.001   | 1.38 (1.14–1.68)           |
| Insomnia severity index score | 0.07 (0.01) | <0.001 | 1.07 (1.04–1.10)           |
| Average minutes outside in sunlight per day—no. | 0.00 (0.00) | 0.021   | 1.00 (0.99–1.00)           |
| Age—yr           | −0.02 (0.01) | 0.012   | 0.98 (0.96–1.00)           |
| Female sex       | −0.54 (0.19) | 0.004   | 0.58 (0.40–0.84)           |
| Someone emotionally close…can talk to daily | −0.63 (0.27) | 0.021   | 0.53 (0.31–0.91)           |
| Touch others less | −0.82 (0.35) | 0.018   | 0.44 (0.22–0.87)           |
| Have enough social support to get through this | −0.89 (0.22) | <0.001 | 0.41 (0.27–0.63)           |
| PSQ-9 suicidal ideation | 0.26† | <0.001 |                       |
| Worry about ability to financially support self/family | 0.55 (0.21) | 0.007   | 1.74 (1.16–2.61)           |
| Alcohol drinks per day | 0.31 (0.10) | 0.002   | 1.37 (1.12–1.67)           |
| Insomnia severity index score | 0.09 (0.01) | <0.001 | 1.09 (1.06–1.13)           |
| Age—yr           | −0.03 (0.01) | 0.004   | 0.97 (0.96–0.99)           |
| Days outside in sunlight per week—no. | −0.15 (0.04) | 0.001   | 0.86 (0.79–0.94)           |
| Female sex       | −0.50 (0.19) | 0.009   | 0.61 (0.42–0.88)           |
| Touch others less | −0.77 (0.36) | 0.032   | 0.46 (0.23–0.93)           |
| Have enough social support to get through this | −0.96 (0.21) | <0.001 | 0.38 (0.25–0.58)           |

† Significant at false discovery rate (FDR) correction, p <0.05; n.s., non-significant; yr, year; no., number.

and tens of millions of Americans found themselves out of work. Consistent with the unemployment data from the first month of the shutdown restrictions, 17.4% of the participants in this study reported having lost their primary job as a direct result of the COVID-19 pandemic. While mental health problems were notably high for the sample as a whole, those who lost their primary job directly as a result of the pandemic consistently showed greater severity across all measures of depression, anxiety, and stress responses. The prevalence of clinically significant mental health problems was 1.5–1.7 times higher among those who reported a COVID-19-related job loss than those who did not report such a loss. Large meta-analyses have shown that mental health problems are about twice as prevalent among individuals who are unemployed than those who are employed (35). Our findings suggest that the difference in the rates of mental health problems between those who did
and did not lose their jobs appears to be very similar to those of prior studies, but perhaps slightly smaller in the magnitude of difference, probably due to elevated rates of general pandemic-related concerns even among those who did not experience a COVID-related job loss. Unemployment and financial insecurity are well known contributors to poorer mental health, including depression and suicide (6). Our data are consistent with the existing literature and suggest that the rise in unemployment during the pandemic is associated with significantly elevated mental health problems.

For the sample as a whole, poorer mental health outcomes tended to be predicted by greater worry about the ability to financially support oneself or loved ones, feeling socially isolated, greater severity of insomnia symptoms, and consuming more alcohol. On the other hand, consistent with prior research, protective factors included spending more days per week outside in the sunlight, perceiving enough social and emotional support to get through the crisis, and older age. The contributory role of each of these factors is not surprising and all have been supported by considerable research (13, 36–43). In particular, numerous studies suggest that younger age groups have exhibited greater mental health problems as a result of the pandemic (40–43). Ways to encourage safe outside activities and facilitate social and emotional support need to be explored and encouraged to help individuals maintain resilience and wellbeing during the pandemic stay-at-home period.

How can these data inform psychiatric care and public health policy? The data suggest that subjective reports of financial worry represented the most consistent and predictive factor associated with meeting criteria for clinically significant mental health problems. Because of the extraordinarily high level of job loss produced as a direct consequence of the pandemic response, these findings suggest that efforts to address the personal financial impacts of the pandemic are going to be pivotal contributors to averting an impending mental health crisis. Social isolation and a sense of insufficient social support each also contributed significantly to mental health problems. This is consistent with other data suggesting that loneliness has increased during the course of the pandemic, and is associated with suicidal ideation and other mental health issues (36, 44, 45). Clearly, any successful psychiatric mitigation strategy will need to address the profound issues surrounding the current reduction in face-to-face human interaction and the widespread experience of loneliness (46). Problems with insomnia were also highly predictive of poorer mental health, suggesting that sleep assessment should be incorporated into routine clinical contacts and behavioral and medical efforts aimed toward facilitating better sleep health should be a priority (47). Further, suicidal ideation was predicted by greater alcohol intake in combination with financial worries. This is particularly concerning, as recent evidence suggests that alcohol purchases, consumption, and dependence behaviors increased dramatically for those under lockdown during the first 6-months of the pandemic (39, 48, 49). For those at risk of suicidal ideation, alcohol intake should be minimized/avoided (5). Finally, spending more days outside in sunlight was frequently a predictor of positive mental health outcomes. Light exposure is important for enhancing mood and maintaining a healthy sleep schedule (50). Even during prolonged stay-at-home mandates, it is recommended that individuals find ways to increase daylight exposure and, when appropriate, to engage safely and responsibly in appropriately socially distanced outdoor activities to maintain mental health and wellbeing.

A small, but interesting finding is also worthy of note. In Table 3, we found that scores on the GAD-7 were lower among individuals who also reported that they lived with someone in the household who had been diagnosed with COVID-19. We interpret this counterintuitive finding as the effect of seeing COVID-19 first-hand, which may have reduced anxiety over the unknown. During the early weeks of the pandemic, not much was known about the virus, which led to much speculation and widespread worry. Since the vast majority of people who contract COVID-19 tend to be asymptomatic or experience only mild illness, the experience of direct exposure to someone who has been diagnosed and potentially recovered may actually have reduced their anxiety by making the illness concrete. Of course, this is post-hoc speculation, and will require further research.

While the present sample was collected to be representative of the general U.S. population, it is important to keep in mind that the data may not be representative of the mental health responses in other areas of the world. COVID-19 has affected every country on the globe, but the response to the pandemic has been different across cultures. For instance, cultures that adhere to tightly to social norms appear to have fared much better with regard to the number of COVID-19 cases and deaths than cultures that adhere much more loosely to such norms (such as the U.S.) (51). Consequently, the willingness to accept the necessity or legitimacy of the government lockdowns and their repercussions on social or occupational functioning may play a role in how job loss may be perceived and how it may affect mental health. The present research does not directly address this issue, but it will be one that is important for further study. In the meantime, the generalizability of these findings to other cultures with different values should be considered as tentative until validated with further research in other countries around the world.

This study was limited by its use of self-report measures and online questionnaires rather than in-person clinical interviews. Future work will involve more extensive clinical interviews and longitudinal data collection to monitor changes during the course of the pandemic. Furthermore, at the time when these data were first collected, there were no readily available validated COVID-19 metrics to assess mental health issues, which is a clear limitation. Since that time, validated metrics such as the Fear of COVID-19 Scale have become available (52), and are recommended for use in future COVID-19-related studies. The present findings are also limited by the fact that most of the large epidemiological samples to which we compared our findings were collected some time ago and may differ somewhat from the demographics of the current sample. Population prevalence rates for various disorders change over time and so it is possible that our findings overestimate the prevalence of these mental health issues. Additionally, the questionnaires we used were generally designed for screening purposes rather than comprehensive psychodiagnostic assessment. Such metrics are often designed for high sensitivity relative to specificity, and may lead to an overestimation of the prevalence of certain disorders (34).
These limitations notwithstanding, the present findings strongly suggest that the U.S. population experienced extraordinary mental health concerns in the first weeks after nationwide pandemic restrictions were enacted (14). It is conceivable that these mental health problems will persist or even increase in the coming months and years as the long-term occupational and personal financial fallout from the pandemic continues to be realized. Large scale efforts to mitigate the effects of financial instability and facilitate social connectedness will be crucial to minimizing the long-term impact of the pandemic on mental health.

DATA AVAILABILITY STATEMENT
The raw data supporting the conclusions of this article will be made available by the authors, without undue reservation.

REFERENCES
1. Bayham J, Fenichel EP. Impact of school closures for COVID-19 on the US health-care workforce and net mortality: a modelling study. Lancet Public Health. (2020) 5:E271–8. doi: 10.1016/S2468-2667(20)30315-8
2. Prem K, Liu Y, Russell TW, Kucharski AJ, Eggo RM, Davies N, et al. The effect of control strategies to reduce social mixing on outcomes of the COVID-19 epidemic in Wuhan, China: a modelling study. Lancet Public Health. (2020) 5:E261–70. doi: 10.1016/S2468-2667(20)30330-5
3. Bump P. Mapping where America has been shut down. The Washington Post (March 18, 2020).
4. Long H: U.S. now has 22 million unemployed, wiping out a decade of job gains. The Washington Post (April 17, 2020).
5. Kerr WC, Kaplan MS, Huguet N, Caetano R, Giesbrecht N, McFarland BH. Economic recession, alcohol, and suicide rates: comparative effects of poverty, foreclosure, and job loss. Am J Prev Med. (2017) 52:469–75. doi: 10.1016/j.amepre.2016.09.021
6. Pompili M, Innamorati M, Di Vittorio C, Baratta S, Masotti V, Badaracco A, et al. Unemployment as a risk factor for completed suicide: a psychological autopsy study. Arch Suicide Res. (2014) 18:181–92. doi: 10.1080/13811118.2013.803449
7. Reger MA, Stanley IH, Joiner TE. Suicide mortality and suicide: a psychological autopsy study. JAMA Psychiatry. (2020) 77:1093–4. doi: 10.1001/jamapsychiatry.2020.1060
8. Xiang YT, Jin Y, Cheung T. Joint International Collaboration to combat mental health challenges during the coronavirus disease 2019 pandemic. JAMA Psychiatry. (2020) 77:989–90. doi: 10.1001/jamapsychiatry.2020.10.20.1057
9. Nie X, Feng K, Wang S, Li Y. Factors influencing public panic during the COVID-19 pandemic. Front Psychol. (2021) 12:576301. doi: 10.3389/fpsyg.2021.576301
10. Dutheil M, Mondillon L, Navel V. PTSD as the second tsunami of the SARS-Cov-2 pandemic. Psychol Med. (2020) 2020:1–2. doi: 10.1017/S0033291720001336
11. Liu N, Zhang F, Wei C, Jia Y, Shang Z, Sun L, et al. Prevalence and predictors of PTSS during COVID-19 outbreak in China hardest-hit areas: gender differences matter. Psychiatry Res. (2020) 287:112921. doi: 10.1016/j.psychres.2020.112921
12. Hawryluck L, Gold WL, Robinson S, Pogorski S, Galea S, Styra R. SARS control and psychological effects of quarantine, Toronto, Canada. Emerg Infect Dis. (2004) 10:1206–12. doi: 10.3201/eid1007.030703
13. Killgore WDS, Taylor EC, Cloonan SA, Dailey NS. Psychological resilience during the COVID-19 lockdown. Psychiatry Res. (2020) 291:113216. doi: 10.1016/j.psychres.2020.113216
14. Pfefferbaum B, North CS. Mental health and the covid-19 pandemic. N Engl J Med. (2020) 383:510–2. doi: 10.1056/NEJMmp2008817
15. Litman L, Robinson J, Abberbock T. TurkPrime.com: a versatile crowdsourcing data acquisition platform for the behavioral sciences. Behav Res Methods. (2017) 49:433–42. doi: 10.3758/s13428-016-0727-7
16. Beck AT, Steer RA, Brown GK. BDHI: Beck Depression Inventory Manual. 2nd ed. San Antonio, TX, Psychological Corporation (1996). doi: 10.1037/100742-000
17. Kroenke K, Spitzer RL, Williams JB. The PHQ-9: validity of a brief depression severity measure. J Gen Intern Med. (2001) 16:606–13. doi: 10.1046/j.1525-1497.2001.016008206.x
18. Spitzer RL, Kroenke K, Williams JB, Lowe B. A brief measure for assessing generalized anxiety disorder: the GAD-7. Arch Intern Med. (2006) 166:1092–7. doi: 10.1001/archinte.166.10.1092
19. Plummer F, Manea L, Trelap D, McMillan D. Screening for anxiety disorders with the GAD-7 and GAD-2: a systematic review and diagnostic metaanalysis. Gen Hosp Psychiatry. (2016) 39:24–31. doi: 10.1016/j.genhospsyc.2015.11.005
20. Zhang WW. A rating instrument for anxiety disorders. Psychosomatics. (1971) 12:371–9. doi: 10.1016/S0033-3182(71)71479-0
21. Dunstan DA, Scott N. Norms for Zhang’s Self-rating Anxiety Scale. BMJ Psychiatry. (2020) 20:90. doi: 10.1136/bmjpsychiatry-2017-102472
22. Spielberger CD, Gorsuch RL, Lushene RE. Manual for the State-Trait Anxiety Inventory. Palo Alto, CA, Consulting Psychologist Press (1970).
23. Fisher PL, Durham RC. Recovery rates in generalized anxiety disorder following psychological therapy: an analysis of clinically significant change in the STAI-T across outcome studies since 1990. Psychol Med. (1999) 29:1425–34. doi: 10.1017/S0033291799001336
24. Bebeau R, Mischel E, Resnick H, Kilpatrick D, Friedman M, Craske M. Dimensional assessment of posttraumatic stress disorder in DSM-5. Psychiatry Res. (2014) 218:143–7. doi: 10.1016/j.psychres.2013.03.032
25. Kilpatrick DG, Resnick HS, Friedman MJ. Severity of Acute Stress Symptoms—Adult: National Stressful Events Survey Acute Stress Disorder Short Scale (NSESSS). American Psychiatric Association (2013).
26. Ouimette P, Wade M, Prins A, Schohn M. Identifying PTSD in primary care: comparison of the Primary Care-PTSD screen (PC-PTSD) and the General Health Questionnaire-12 (GHQ). J Anxiety Disorders. (2008) 22:337–43. doi: 10.1016/j.janxdis.2007.02.010
27. Frings P, Ouimette P, Kimerling R, Camerond RP, Guelshofer DS, Shaw-Hegwer J, et al. The primary care PTSD screen (PC-PTSD): development and operating characteristics. Primary Care Psychiatry. (2004) 9:9–14. doi: 10.1185/135525703215002360
28. Bastien CH, Vallieres A, Morin CM. Validation of the Insomnia Severity Index as an outcome measure for insomnia research. Sleep Med. (2001) 2:297–307. doi: 10.1016/S1389-9457(00)00065-4

ETHICS STATEMENT
The studies involving human participants were reviewed and approved by University of Arizona Institutional Review Board. The patients/participants provided their electronic informed consent to participate in this study.

AUTHOR CONTRIBUTIONS
WK primary study design and conceptualization, primary literature search, data analysis, data interpretation, writing of the initial draft, and figures and tables. SC contributed to study design, data collection, and editing drafts of manuscript. ET and ND contributed to study design and editing drafts of manuscript. All authors contributed to the article and approved the submitted version.
29. Morin CM, Belleville G, Belanger L, Ivers H. The Insomnia Severity Index: psychometric indicators to detect insomnia cases and evaluate treatment response. Sleep. (2011) 34:601–8. doi: 10.1093/sleep/34.5.601

30. Kessler RC, Petukhova M, Sampson NA, Zaslavsky AM, Wittchen HU. Twelve-month and lifetime prevalence and lifetime morbidity risk of anxiety and mood disorders in the United States. Int J Methods Psychiatric Res. (2012) 21:69–84. doi: 10.1002/mpr.1359

31. Crosby AE, Han B, Ortega LA, Parks SE, Groer J. Suicidal thoughts and behaviors among adults aged ≥18 years—United States, 2008–2009. MMWR Surveill Summ. (2011) 60:1–22. Available online at: https://www.cdc.gov/mmwr/pdf/ss/ss6013a1.htm

32. Heinze G, Wallisch C, Dunkler D. Variable selection - a review and recommendations for the practicing statistician. Biom J. (2018) 60:431–49. doi: 10.1002/bimj.201700067

33. Ettman CK, Abdalla SM, Cohen GH, Sampson L, Vivier PM, Galea S. Prevalence of depression symptoms in US ADULTS BEFORE AND DURING the COVID-19 pandemic. JAMA Netw Open. (2020) 3:e2019686. doi: 10.1001/jamanetworkopen.2020.19686

34. Thombs BD, Kwakkenbos L, Levis AW, Benedetti A. Addressing overestimation of the prevalence of depression based on self-report screening questionnaires. Can Med Assoc J. (2018) 190:E44–9. doi: 10.1503/cmaj.1706911

35. Paul KI, Moser K. Unemployment impairs mental health: meta-analyses. J Vocat Behav. (2009) 74:264–82. doi: 10.1016/j.jvb.2009.01.001

36. Killgore WDS, Cloonan SA, Taylor EC, Dailey NS. Loneliness: a signature mental health concern in the era of COVID-19. Psychiatry Res. (2020) 290:113117. doi: 10.1016/j.psychres.2020.113117

37. Killgore WDS, Cloonan SA, Taylor EC, Fernandez F, Grandner MA, Dailey NS. Suicidal ideation during the COVID-19 pandemic: the role of insomnia. Psychiatry Res. (2020) 290:113134. doi: 10.1016/j.psychres.2020.113134

38. Li F, Luo S, Mu W, Li Y, Ye L, Zheng X, et al. Effects of sources of social support overestimation of the prevalence of depression based on self-report screening questionnaire. J Psychiatr Res. (2021) 195:10.1016/j.jpsychires.2021.01.047

39. Killgore WDS, Cloonan SA, Taylor EC, Miller MA, Dailey NS. Loneliness during the first half-year of COVID-19 lockdowns. Psychiatry Res. (2020) 293:113492. doi: 10.1016/j.psychres.2020.113492

40. El-Gabalawy R, Sommer JL. Are we at risk too: the disparate mental health impacts of the pandemic on younger generations: nous sommes aussi a risque: les effets disparates de la pandemie sur la sante mentale des generations plus jeunes. Can J Psychiatry. (2021) 2021:706743721989162. doi: 10.1177/0706743721989162

41. Novotny JS, Gonzalez-Rivas JP, Kunzova S, Skladana M, Pospisilova A, Polcrova A, et al. Risk factors underlying COVID-19 lockdown-induced mental distress. Front Psychiatry. (2020) 11:603014. doi: 10.3389/fpsyg.2020.603014

42. Wang Y, Kala MP, Jafar TH. Factors associated with psychological distress during the coronavirus disease 2019 (COVID-19) pandemic on the predominantly general population: a systematic review and meta-analysis. PLoS ONE. (2020) 15:e0244630. doi: 10.1371/journal.pone.0244630

43. Wickens CM, McDonald AJ, Elton-Marshall T, Wells S, Nigatu YT, Jankowicz D, et al. Loneliness in the COVID-19 pandemic: associations with age, gender and their interaction. J Psychiatr Res. (2021) 136:103–8. doi: 10.1016/j.jpsychires.2021.01.047

44. Killgore WDS, Cloonan SA, Taylor EC, Miller MA, Dailey NS. Three months of loneliness during the COVID-19 lockdown. Psychiatry Res. (2020) 293:113392. doi: 10.1016/j.psychres.2020.113392

45. Killgore WDS, Cloonan SA, Taylor EC, Lucas DA, Dailey NS. Loneliness and resilience on the mental health of different age groups during the COVID-19 pandemic. Int J Ment Health Addict. (2021) 2021:1–4. doi: 10.1080/08897077.2020.1856293

46. Pail G, Hufl W, Pyrek E, Winkler D, Willeit M, Praschak-Rieder N, et al. Bright-light therapy in the treatment of mood disorders. Neuropsychobiology. (2011) 64:152–62. doi: 10.1159/000328950

47. McKetta S, Morrison CN, Keyes KM. Trends in US alcohol consumption frequency during the first wave of the SARS-CoV-2 pandemic. Alcohol Clin Exp Res. (2021) 2021:10.1111/acer.14575. doi: 10.1111/acer.14575

48. Morton CM. Alcohol sales during COVID-19 social restrictions: initial evidence from alcoholic beverage control states. Subst Abus. (2021) 3:1–4. doi: 10.1080/08897077.2020.1856293

49. Gelfand MJ, Jackson JC, Pan X, Nau D, Pieper D, Denison E, et al. The relationship between cultural tightness–looseness and COVID-19 cases and deaths: a global analysis. Lancet Psychiatry. (2021) 8:479–58. doi: 10.1016/S2215-0366(17)30328-0

50. Ahorsu DK, Lin CY, Imani V, Saffari M, Griffiths MD, Pakpour AH. The fear of covid-19 scale: development and initial validation. Int J Ment Health Addict. (2020) 2020:1–9. doi: 10.1007/s11469-020-00270-8

Conflict of Interest: The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

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