SCALING LIMITS OF COUPLED CONTINUOUS TIME RANDOM WALKS AND RESIDUAL ORDER STATISTICS THROUGH MARKED POINT PROCESSES

ADAM BARCZYK AND PETER KERN

Abstract. A continuous time random walk (CTRW) is a random walk in which both spatial changes represented by jumps and waiting times between the jumps are random. The CTRW is coupled if a jump and its preceding or following waiting time are dependent random variables (r.v.), respectively. The aim of this paper is to explain the occurrence of different limit processes for CTRWs with forward- or backward-coupling in Straka and Henry (2011) using marked point processes. We also establish a series representation for the different limits. The methods used also allow us to solve an open problem concerning residual order statistics by LePage (1981).

1. Introduction

Two i.i.d. sequences of $\mathbb{R}^+$-valued waiting times $(J_n)_{n \in \mathbb{N}}$ and of $\mathbb{R}^d$-valued jumps $(X_n)_{n \in \mathbb{N}}$ yield two versions of a CTRW by

$$S_{N_t} := \sum_{k=1}^{N_t} X_k, \quad S_{N_{t+1}} := \sum_{k=1}^{N_{t+1}} X_k,$$

where $N_t := \max\{n \in \mathbb{N}_0 : \sum_{k=1}^n J_k \leq t\}$ is the number of jumps up to time $t$. The CTRW is coupled if the sequences $(J_n)_{n \in \mathbb{N}}$ and $(X_n)_{n \in \mathbb{N}}$ are dependent. Typically we assume, that the sequence $(J_n, X_n)_{n \in \mathbb{N}}$ is i.i.d. with unknown dependence between the waiting time $J_n$ and the jump $X_n$ for fixed $n \in \mathbb{N}$. Using this dependence structure $S_{N_t}$ is called backward-coupled CTRW whereas $S_{N_{t+1}}$ is called forward-coupled CTRW. Both processes represent the position of a jumper at time $t$, but in the backward-coupled case the particle first waits for a time $J_1$ before jumping to $X_1$, whereas in the forward-coupled case the particle jumps to $X_1$ at time $t = 0$ and then waits for a
time \( J_1 \) and so on. CTRW processes were introduced in [22] to study random walks on a lattice and have been studied intensively over the past few decades. Today there is a wide field of possible applications for CTRWs. They are used in physics to model phenomena of anomalous diffusion [28]. The jumps can also represent movements of an ensemble of particles being transported over the earth surface in geophysics [27] or represent log-returns in finance [25]. A comprehensive study of limit theorems for coupled CTRWs has been initiated in [3] covering previously known special models [28, 14, 15] from physics. The limiting distributions of forward- and backward-coupled CTRWs have been investigated by Straka and Henry [29] using a continuous mapping approach on the space of their sample paths. A similar approach in a more general setting appears in [11]. Straka and Henry prove that the limiting processes of coupled CTRWs in general differ when waiting times precede or follow jumps, respectively. Also the differences between the properties of these processes are not marginal, cf. [12]. Unfortunately neither the continuous mapping approach used in [29] nor the methods used in [12] are adequate to point out why different scaling limits occur. So a new approach to fill this gap is made using marked point processes here. Defining the time of the \( n \)-th jump by \( T_n := \sum_{k=1}^{n} J_k \) we first study the limit behavior of the point processes which arise by marking each jump time with its occurring jump, respectively, i.e. we analyse

\[
\sum_{k=1}^{n} \varepsilon(T_k, X_k), \quad \sum_{k=1}^{n} \varepsilon(T_{k-1}, X_k),
\]

(1.1)

It turns out that only the jumps with large norm contribute to the limit distributions of (1.1), as it is already known for real-valued partial sums which converge to an infinitely divisible r.v. without Gaussian part, cf. [1] and references therein. The methods used also solve an open problem concerning the convergence of residual order statistics by LePage, cf. [17], [18], [26]. The scaling limits of the CTRWs can be determined by summing up the marks of the points in (1.1) which have a jump time occurring before time \( t \). Hence in the scaling limit of forward-coupled CTRWs an additional big jump occurs compared to its backward-coupled version, which illuminates the difference between the processes. This approach also provides a series representation for the different scaling limits which might be of interest for
simulation purposes. Since the resulting limit processes are not Lévy processes, no efficient simulation algorithm is known yet.

2. Preliminaries

Let \((J_n, X_n)_{n \in \mathbb{N}}\) be an i.i.d. sequence of \(\mathbb{R}^+ \times \mathbb{R}^d\)-valued r.v. Assume that \((J_1, X_1)\) belongs to the generalized domain of attraction (GDOA) of a r.v. \((D, A)\), where \(D\) is stable with index \(\alpha \in (0, 1)\) and \(A\) is full operator stable with index \(E \in GL(\mathbb{R}^d)\) and without Gaussian part. Note that by Theorem 7.2.1 of [20] the real parts of the eigenvalues of \(E\) are greater than \(1/2\). By classic results [13, Theorem 14.14] and [21, Theorem 4.1] this implies that there exist regularly varying sequences \((b_n)_{n \in \mathbb{N}} \in RV_{1/\alpha}\) and \((A_n)_{n \in \mathbb{N}} \in RV_{-E}\) such that the convergence

\[
\left( b_n^{-1} \sum_{k=1}^{\lfloor nt \rfloor} J_k, \sum_{k=1}^{\lfloor nt \rfloor} (A_n X_k - \mathbb{E}(A_n X_k 1_{\|A_n X_k\| \leq \tau})) \right) \xrightarrow{D} (D(t), A(t))_{t \geq 0}
\]

holds in \(D([0, \infty), \mathbb{R}^+ \times \mathbb{R}^d)\) for any \(\tau > 0\) such that for the Lévy measure \(\eta\) of \(A\) and the sphere \(S^d_{\tau} = \{ x \in \mathbb{R}^d : \| x \| = \tau \}\) we have \(\eta(S^d_{\tau}) = 0\), i.e. the sphere \(S^d_{\tau}\) is a continuity set for \(\eta\). Here the process \(D(\cdot)\) denotes an \(\alpha\)-stable subordinator and \(A(\cdot)\) denotes an operator Lévy motion. Note that the drift term of the Lévy process \(A\) depends on \(\tau\). It is well known that we can choose \(\tau = \infty\) if the real part of any eigenvalue of \(E\) belongs to \((1/2, 1)\), since then \(\mathbb{E}(X_1)\) exists. Moreover we can choose \(\tau = 0\) if the real part of any eigenvalue of \(E\) exceeds 1. Due to the spectral decomposition in [20], centering by truncated expectations in (2.1) is only necessary if some eigenvalue of the exponent \(E\) has real part equal to 1.

We already stated, that only points with large norm contribute to the limit of the point processes in (1.1). So we use a radial decomposition of the Lévy measure

\[
\eta(A) = \int_{S^{d-1}} \int_0^\infty 1_A(xv) \tilde{\eta}(dx, v) d\sigma(v)
\]

where \(\sigma\) is a probability measure on the unit sphere \(S^{d-1}\) of \(\mathbb{R}^d\) and \((\tilde{\eta}(\cdot, v))_{v \in S^{d-1}}\) is a weakly measurable family of Lévy measures on \((0, \infty)\), cf. [24]. We also define the right-continuous inverse of \(\tilde{\eta}(\cdot, v)\) by

\[
\tilde{\eta}^+(x, v) := \sup\{ u > 0 : \tilde{\eta}([u, \infty), v) \geq x \}.
\]
With this notation we are able to give a series representation for the process \( A(\cdot) \) in \( D([0, T], \mathbb{R}^d) \) for fixed \( T > 0 \) by

\[
\lim_{\varepsilon \downarrow 0} \left( \sum_{T - \tau_k \leq t} (\tilde{\eta}^{-1}(T^{-1} \Gamma_k, \mathbf{V}_k) \mathbf{V}_k 1_{\eta^{-1}(T^{-1} \Gamma_k, \mathbf{V}_k) \geq \varepsilon}) - t \int_{\varepsilon \leq \|x\| \leq \tau} x \, d\eta(x) \right),
\]

cf. \([17], [24]\), where \( \Gamma_n \) is the \( n \)-th partial sum of i.i.d. standard exponential r.v., \( (\tau_n)_{n \in \mathbb{N}} \) denotes an i.i.d. sequence of uniformly \( U(0, 1) \)-distributed r.v. and \( (\mathbf{V}_n)_{n \in \mathbb{N}} \) denotes an i.i.d. sequence with distribution \( \sigma \), with \( (\Gamma_n)_{n \in \mathbb{N}}, (\tau_n)_{n \in \mathbb{N}} \) and \( (\mathbf{V}_n)_{n \in \mathbb{N}} \) being independent.

Now it is well known that for a triangular array of infinitesimal row-wise independent \( \mathbb{R}^+ \)-valued r.v. \((Y_{k,n})_{1 \leq k \leq n}, n \in \mathbb{N}\), converging to an infinitely divisible r.v. \( Y \) with associated Lévy measure \( \phi \), only the extremes contribute to the limit distribution, cf. \([1]\) and references therein. This result coincides with the convergence

\[
\sum_{k=1}^{n} \varepsilon \left( \frac{k}{n}, Y_{k,n} \right) \xrightarrow{D} \text{PRM} (\lambda \otimes \phi)
\]
in \( M_p([0, 1] \times (0, \infty]) \), where \( \text{PRM} (\lambda \otimes \phi) \) denotes a Poisson random measure with mean measure \( \lambda \otimes \phi \). Furthermore, it is well known that \( \sum_{k \in \mathbb{N}} \varepsilon(\tau_k, \phi^{-1}(\Gamma_k)) \) is also a representation of \( \text{PRM} (\lambda \otimes \phi) \) in \( M_p([0, 1] \times (0, \infty]) \), where \( \phi^{-1} \) denotes the right-sided inverse of \( \phi \), cf. \([24]\). This fact can be understood by sorting the points on the left-hand side in \((2.4)\)

\[
\sum_{k=1}^{n} \varepsilon \left( \frac{k}{n}, Y_{k,n} \right) = \sum_{k=1}^{n} \varepsilon \left( \frac{d_k}{n}, Y_{n-k+1:n} \right),
\]

where \( (Y_{1:n}, \ldots, Y_{n:n}) \) denotes the order statistics of \((Y_{1,n}, \ldots, Y_{n,n})\) with corresponding antirank vector \((d_1, \ldots, d_n)\), i.e. the inverse permutation of the rank vector. Using Freedman’s Lemma, cf. \([10]\), one can easily verify, that the convergence \((n^{-1}d_k)_{k \in \mathbb{N}} \xrightarrow{D} (\tau_k)_{k \in \mathbb{N}} \) holds in \([0, 1]^N\), where \( d_k = 0 \) for \( k > n \). Moreover, \((2.4)\) in \([1]\) gives us \( Y_{n-k+1:n} \xrightarrow{D} \phi^{-1}(\Gamma_k) \). So the convergence in \((2.4)\) can also be established by analysing the convergence of the points \((n^{-1}d_k, Y_{n-k+1:n})_{1 \leq k \leq n}\). This approach can also be applied to the point processes in \((1.1)\). As the r.v. \((X_n)_{n \in \mathbb{N}}\) are \( \mathbb{R}^d \)-valued one cannot use traditional order statistics. LePage suggested to use a normwise sorting, cf. \([17]\). So for \( x_1, \ldots, x_n \in \mathbb{R}^d \) we introduce the residual order statistics \( x_{1:n}, \ldots, x_{n:n} \) by \( \|x_{1:n}\| \leq \ldots \leq \|x_{n:n}\| \).
3. Convergence of residual order statistics

The convergence of the normalized residual order statistics $A_n X_{n-k+1:n}$ is still an open problem. LePage [17] conjectures, that a generalisation of the one-dimensional case

$$
(A_n X_{n-k+1:n})_{k \in \mathbb{N}} \xrightarrow{D} (\bar{\eta}^\leftarrow (\Gamma_k, V_k) V_k)_{k \in \mathbb{N}}
$$

holds. As usual one sets $X_{k:n} = 0$, whenever $k \leq 0$ or $k > n$. This result has been proven in [18] for the case that the limit process $A(\cdot)$ is multivariate $\alpha$-stable. In this case the right-sided inverse $\bar{\eta}^\leftarrow (x, v)$ defined in (2.2) is independent of $v \in \mathbb{S}^{d-1}$ as the projection of the Lévy measure $\eta$ is the same for every direction, cf. Theorem 7.3.3 in [20]. Some years later a similar problem has been studied in [9] using a different norm $\| \cdot \|_H$ which respects the special structure of the operator $E$. In [26] the operator semistable case has been studied, but the result (3.1) also could only be established in the special case, that $\bar{\eta}^\leftarrow (x, v)$ is independent of $v$, which concides with the multivariate $\alpha$-stable case. The author also supposed that the convergence (3.1) holds only in this case. We will show that the limit on the right-hand side in (3.1) has to be modified. The proof is based on the following lemma.

**Lemma 1.** Let $N_n = \sum_{k=1}^n \varepsilon X_k^{(n)}$, $n \in \mathbb{N}_0$, be a sequence of point processes in $M_p([-\infty, \infty]^d \setminus \mathbb{K}_\varepsilon^d)$, the set of all point measures on the space $[-\infty, \infty]^d \setminus \mathbb{K}_\varepsilon^d$, where $\mathbb{K}_\varepsilon^d := \{x \in \mathbb{R}^d : \|x\| \leq \varepsilon\}$ denotes the compact $\varepsilon$-ball in $\mathbb{R}^d$. Suppose $N_n \xrightarrow{D} N_0$. If $\varepsilon < \|X_1^{(0)}\| < \infty$ holds for every $i \in \mathbb{N}$ almost surely (a.s.) and $\|X_j^{(0)}(\omega)\| > \|X_i^{(0)}(\omega)\|$ for all $1 \leq i < j$ a.s. then the convergence

$$
(A_n X_{n-k+1:n})_{k \in \mathbb{N}} \xrightarrow{D} \left(X_k^{(0)}\right)_{k \in \mathbb{N}}
$$

holds in $(\mathbb{R}^d \setminus \mathbb{K}_\varepsilon^d)^N$.

**Proof.** The proof is based on a continuous mapping approach and Lemma 7.1 in [23]. Define $M \subset M_p([-\infty, \infty]^d \setminus \mathbb{K}_\varepsilon^d)$ by

$$
M := \left\{m : m = \sum_{k=1}^P \varepsilon x_k, \infty > \|x_1\| > \ldots > \|x_P\| > \varepsilon\right\}.
$$

Now we show that the mapping

$$
\pi_k : M_p([-\infty, \infty]^d \setminus \mathbb{K}_\varepsilon^d) \to [-\infty, \infty]^d \setminus \mathbb{K}_\varepsilon^d
$$

...
\[
\pi_k \left( \sum_{i=1}^{P} \varepsilon_{x_i} \right) \mapsto x_{P-k+1:P}, \ x_{P-k+1:P} = 0 \text{ for } k \leq 0 \text{ and } k > P
\]

is continuous in \( m \in M \) for every \( k \in \mathbb{N} \). Let \((m_n)_{n \in \mathbb{N}} \subset M_p([-\infty, \infty]^d \setminus \mathbb{K}_\varepsilon^d)\) be a sequence of point measures converging vaguely to a point measure \( m_0 = \sum_{k=1}^{P} \varepsilon_{x_k} \in M_p([-\infty, \infty]^d \setminus \mathbb{K}_\varepsilon^d)\). Now choose \( n \) sufficiently large so that all points of \( m_n \) lie inside of \([-\infty, \infty]^d \setminus \mathbb{K}_\varepsilon^d\). By sorting the points of \( m_n \) in descending order of their norm

\[
m_n = \sum_{i=1}^{P} \varepsilon_{x_i(n)}, \ \varepsilon < \|x_1^{(n)}\| \leq \ldots \leq \|x_P^{(n)}\| < \infty
\]

an application of Lemma 7.1 in [23] yields the convergence of the points

\[
(3.2) \quad \left( x_1^{(n)}, \ldots, x_P^{(n)} \right) \longrightarrow \left( x_1^{(0)}, \ldots, x_P^{(0)} \right)
\]

in \((\mathbb{R}^d \setminus \mathbb{K}_\varepsilon^d)^P\). Now by the definition of the mapping \( \pi_k \)

\[
\pi_k(m_n) = x_k^{(n)}, \ \pi_k(m_0) = x_k^{(0)} \text{ for } 1 \leq k \leq P
\]

\[
\pi_k(m_n) = \pi_k(m_0) = 0 \text{ for } k > P
\]

holds and \( \pi_k \) is continuous by (3.2) for every \( k \in \mathbb{N} \). An easy application of the continuous mapping theorem yields the desired result. \( \square \)

Lemma 1 allows to identify the distribution of the limit of properly normalized residual order statistics.

**Theorem 2.** For \( k \in \mathbb{N}, \ T > 0 \) and \( \omega \in \Omega \) define

\[
(3.3) \quad \hat{d}_k(\omega) = \arg \left( \max_{i \in \mathbb{N} \setminus \{d_1(\omega), \ldots, d_k(\omega)\}} \eta^-(T^{-1} \Gamma_i(\omega), \mathbf{V}_i(\omega)) \right)
\]

as the argument of the \( k \)-th largest element of the set \( \{\eta^-(T^{-1} \Gamma_i(\omega), \mathbf{V}_i(\omega)), i \in \mathbb{N}\} \).

Then convergence of the residual order statistics

\[
(3.4) \quad \left( A_n \mathbf{X}_{[nT]-k+1:nT]} \right)_{k \in \mathbb{N}} \overset{\mathbb{P}}{\longrightarrow} \left( \eta^-(T^{-1} \Gamma_{\hat{d}_k(\omega)}, \mathbf{V}_{\hat{d}_k(\omega)}) \mathbf{V}_{\hat{d}_k(\omega)} \right)
\]

holds in \((\mathbb{R}^d)^N\).

**Remark 3.** Note that \( \hat{d}_k \) is well-defined for all \( k \in \mathbb{N} \), since the number of elements in the set \( \{i \in \mathbb{N} : \eta^-(T^{-1} \Gamma_i, \mathbf{V}_i) > \varepsilon\} \) is finite a.s. for all \( \varepsilon > 0 \). Moreover theorem 2 does not contradict any of the results proven in [9], [18], [26]. If \( \mathbf{A}(\cdot) \) is a
multivariate $\alpha$-stable Lévy process, the monotonicity of the mapping $x \mapsto \tilde{\eta}^\varepsilon(x, v)$ yields $\hat{d}_k(\omega) = k$ a.s. for all $k \in \mathbb{N}$.

**Proof of Theorem** 2. First we have to determine the limit of the truncated point process

$$\sum_{k=1}^{[nT]} \xi(A_nX_k 1_{\|A_nX_k\| \geq \varepsilon}),$$

where $\varepsilon > 0$ has to be chosen such that $\eta(S^{d-1}_\varepsilon) = 0$ holds. By Theorem 3.2.2 in [20] assumption (2.1) yields the vague convergence

$$[nT] \mathbb{P}(A_nX_k \in \cdot) \xrightarrow{\nu} T \cdot \eta(\cdot)$$

in $\mathbb{R}^d \setminus \{0\}$. Hence the convergence of the point processes

$$\sum_{k=1}^{[nT]} \xi A_nX_k \xrightarrow{D} \text{PRM}(T \cdot \eta)$$

holds in $M_p([-\infty, \infty]^d \setminus \{0\})$. Now by [21] $\sum_{k \in \mathbb{N}} \xi \tilde{\eta}^{-(T-1)\Gamma_k, V_k} V_k 1_{\tilde{\eta}^{-(T-1)\Gamma_k, V_k} \geq \varepsilon} = \sum_{k \in \mathbb{N}} \xi \tilde{\eta}^{-(T-1)\Gamma_k, V_k} \xi V_k 1_{\tilde{\eta}^{-(T-1)\Gamma_k, V_k} \geq \varepsilon}$. Applying the a.s. continuous restriction functional

$$\pi': M_p([0, \infty]^d \setminus \{0\}) \to M_p([0, \infty]^d), m \mapsto m|_{(\mathbb{K}^d)^c},$$

the continuous mapping theorem yields

$$\sum_{k=1}^{[nT]} \xi A_nX_k 1_{\|A_nX_k\| \geq \varepsilon} \xrightarrow{D} \sum_{k \in \mathbb{N}} \xi \tilde{\eta}^{-(T-1)\Gamma_k, V_k} V_k 1_{\tilde{\eta}^{-(T-1)\Gamma_k, V_k} \geq \varepsilon}.$$

The continuity of $\pi'$ is proven in [5] for instance. Now the points of the point process on the right-hand side of (3.7) have to be ordered in descending order of their norm

$$\sum_{k \in \mathbb{N}} \xi \tilde{\eta}^{-(T-1)\Gamma_k, V_k} V_k 1_{\tilde{\eta}^{-(T-1)\Gamma_k, V_k} \geq \varepsilon} = \sum_{k \in \mathbb{N}} \xi \tilde{\eta}^{-(T-1)\Gamma_k, V_k} \xi V_k 1_{\tilde{\eta}^{-(T-1)\Gamma_k, V_k} \geq \varepsilon}.$$

An application of Lemma [1] yields the convergence of the points

$$\left(A_nX_{[nT]-k+1:[nT]} 1_{\|A_nX_{[nT]-k+1:[nT]}\| \geq \varepsilon}\right)_{k \in \mathbb{N}} \xrightarrow{D} \left(\tilde{\eta}^{-(T-1)\Gamma_k, V_k} V_k 1_{\tilde{\eta}^{-(T-1)\Gamma_k, V_k} \geq \varepsilon}\right)_{k \in \mathbb{N}}$$

in $(\mathbb{R}^d \setminus \mathbb{K}^d)^\mathbb{N}$. The desired result follows by taking the limit as $\varepsilon \downarrow 0$ and an easy application of Theorem 4.2 in [4]. □
4. Convergence of associated point processes

Now that the limit distribution of normalized residual order statistics is identified we can study the associated marked point processes

\[
\sum_{k=1}^{\lfloor nT \rfloor} \mathcal{E}(b_n^{-1} T_k, A_n X_k).
\]

In the uncoupled case convergence results for this processes can be established with a continuous mapping approach using the time deformation defined in [23, (8.29)]. Since the continuity of this time deformation demands the processes \( A(\cdot) \) and \( D(\cdot) \) to have a.s. no common jumps, which is not necessarily fulfilled in the coupled case, this standard methods cannot be applied in our case. So we use a sorting argument like in (2.5).

Lemma 4. Let \((\tau_n)_{n \in \mathbb{N}}, (\Gamma_n)_{n \in \mathbb{N}} \) and \((V_n)_{n \in \mathbb{N}}\) be as in (2.3). Then the convergence of the associated point processes

\[
\begin{align*}
\sum_{k=1}^{\lfloor nT \rfloor} \mathcal{E}(b_n^{-1} T_k, A_n X_k) &\xrightarrow{D} \sum_{k \in \mathbb{N}} \mathcal{E}(D(T^{-1} \tau_k), \tilde{\eta}^-(T^{-1} \Gamma_k, V_k)V_k) \\
\sum_{k=1}^{\lfloor nT \rfloor} \mathcal{E}(b_n^{-1} T_k-1, A_n X_k) &\xrightarrow{D} \sum_{k \in \mathbb{N}} \mathcal{E}(D(T^{-1} \tau_k-1), \tilde{\eta}^-(T^{-1} \Gamma_k, V_k)V_k)
\end{align*}
\]

holds in \( M_p([0, \infty) \times [-\infty, \infty]^d \setminus \{0\}) \) for every \( T > 0 \), where \( D(x-) \) denotes the left-hand limit of the process \( D(\cdot) \) in \( x \).

Proof. Choose \( T > 0 \) arbitrary. We start by sorting the points of the associated point process

\[
\left(\sum_{k=1}^{\lfloor nT \rfloor} \mathcal{E}(b_n^{-1} T_k, A_n X_k)\right) = \left(\sum_{k=1}^{\lfloor nT \rfloor} \mathcal{E}(b_n^{-1} T_{d_{\lfloor nT \rfloor}^{-k+1}}, A_n X_{\lfloor nT \rfloor^{-k+1}})\right).
\]

Again \((d_1, \ldots, d_{\lfloor nT \rfloor})\) denotes the antirank vector of the r.v. \((X_1, \ldots, X_{\lfloor nT \rfloor})\). The normalized residual order statistics \( A_n X_{\lfloor nT \rfloor^{-k+1}} \) have already been studied in Theorem 2. It remains to determine the limit distribution of

\[
b_n^{-1} T_{d_{\lfloor nT \rfloor}^{-k+1}} = b_n^{-1} \sum_{l=1}^{d_{\lfloor nT \rfloor}^{-k+1}} J_l = b_n^{-1} \sum_{l=1}^{\lfloor nT \rfloor} J_l.
\]
Since $X_1, \ldots, X_{[nT]}$ are i.i.d., $(n^{-1}d_k)_{k \in \mathbb{N}} \xrightarrow{D} (T \cdot \tau_k)_{k \in \mathbb{N}}$ suggests that the convergence
\begin{equation}
(b_n^{-1}T_{d_{[nT]-k+1}})_{k \in \mathbb{N}} \xrightarrow{D} (D(T \cdot \tau_k))_{k \in \mathbb{N}}
\end{equation}
holds in $[0, \infty)^N$. But this result cannot be established with a traditional continuous mapping approach. The mapping $\pi_t : D([0, \infty), \mathbb{R}^d) \times \mathbb{R}^+ \rightarrow \mathbb{R}$, $(x, t) \mapsto x(t)$ is only a.s. continuous if $x$ is a.s. continuous in $t$. Also classical transfer theorems, cf. [7, 8], are not helpful because they require independence of the summands and their none of these conditions is fulfilled another approach is used.

Let $(\hat{d}_1, \ldots, \hat{d}_{[nT]})$ denote the associated antirank vector of the waiting times $(J_1, \ldots, J_{[nT]})$. Since the joint convergence of the well-centered and normalized sequential partial sums to the process $(D(\cdot), A(\cdot))$ holds, one can easily prove convergence of the normalized antirank vector
\begin{equation}
(n^{-1}(d_{[nT]-k+1})_{k \in \mathbb{N}}, n^{-1}(\hat{d}_{[nT]-k+1})_{k \in \mathbb{N}}) \xrightarrow{D} \left(\left((T \cdot \tau_i)_{i \in \mathbb{N}}, (T \cdot \hat{\tau}_i)_{i \in \mathbb{N}}\right)\right)
\end{equation}
where $(\tilde{\tau}_n)_{n \in \mathbb{N}}$, $(\hat{\tau}_n)_{n \in \mathbb{N}}$ denote two i.i.d. sequence of $\mathcal{U}(0,1)$-distributed r.v.. Note that the sequences $(\tilde{\tau}_n)_{n \in \mathbb{N}}$ and $(\hat{\tau}_n)_{n \in \mathbb{N}}$ are not independent in the coupled case. As a consequence of the convergence of the antirank vector, the convergence of indicator functions
\begin{equation}
\left(1_{n^{-1}\hat{d}_{[nT]-j+1} \leq n^{-1}d_{[nT]-i+1}}\right)_{j \in \mathbb{N}} \xrightarrow{D} \left(1_{\hat{\tau}_j \leq \tau_i}\right)_{j \in \mathbb{N}}
\end{equation}
holds. An application of Basu’s lemma, cf. [16 Theorem 5.1.2], yields the independence
\begin{equation}(d_1, \ldots, d_{[nT]}), (\hat{d}_1, \ldots, \hat{d}_{[nT]}) \perp (J_{1:[nT]}, \ldots, J_{[nT]:[nT]})\end{equation}
for every fixed $n \in \mathbb{N}$, which proves that the joint convergence
\begin{equation}
(b_n^{-1}J_{[nT]-j+1:[nT]}1_{d_{[nT]-j+1} \leq d_{[nT]-i+1}})_{j \in \mathbb{N}} \xrightarrow{D} \left((\eta^\alpha)^{-1}(T^{-1}\hat{\Gamma}_j)1_{\hat{\tau}_j \leq \tau_i}\right)_{j \in \mathbb{N}}
\end{equation}
holds, where $(\eta^\alpha)^{-1}$ is the right-sided inverse of the Lévy measure associated with $D(1)$ and $(\hat{\Gamma}_n)_{n \in \mathbb{N}}$ denotes a distributional copy of the sequence $(\Gamma_n)_{n \in \mathbb{N}}$. Note that the sequences $(\Gamma_n)_{n \in \mathbb{N}}$ and $(\hat{\Gamma}_n)_{n \in \mathbb{N}}$ are also not independent in the coupled case. Now summation verifies
\begin{equation}
\sum_{l \in \mathbb{N}} b_n^{-1}J_{[nT]-l+1:[nT]}1_{n^{-1}d_{[nT]-l+1} \leq n^{-1}d_{[nT]-i+1}}
\end{equation}
tifies the right-hand side in \((4.6)\) as series representation of convergence of antiranks holds simultaneously, we have proven \((4.6)\) holds. Since the mapping \(\tilde{R}\) in \((4.7)\) yields convergence in \((2.4)\) towards \((2.3)\) can be proven with the same technique, the sequences in \((4.2)\) in \([4]\) yields

\[
\tilde{R} = \sum_{l=1}^{\lfloor nT \rfloor} J_l \mathbf{1}_{l \leq d_{[nT]} - 1 + 1} = b_n^{-1} \sum_{l=1}^{\lfloor d_{[nT]} - 1 + 1 \rfloor} J_l \mathbf{1}_{l \leq d_{[nT]} - 1 + 1} = b_n^{-1} \sum_{l=1}^{\lfloor d_{[nT]} - 1 + 1 \rfloor} J_l \xrightarrow{\mathcal{D}} \sum_{i \in \mathbb{N}} (\eta^\alpha - 1)(T - 1) \tilde{\Gamma}_i \mathbf{1}_{T \tilde{\tau}_i \leq T \tilde{\tau}_i}. \tag{4.6}
\]

Using the Ferguson-Klass series representation of the process \(\tilde{D}(\cdot)\), cf. \([3]\), one identifies the right-hand side in \((4.6)\) as series representation of \(\tilde{D}(T \cdot \tilde{\tau}_i)\). Since the convergence of antiranks holds simultaneously, we have proven

\[
\left( b_n^{-1} \sum_{l=1}^{\lfloor d_{[nT]} - 1 + 1 \rfloor} J_l \right)_{i \in \mathbb{N}} \xrightarrow{\mathcal{D}} (\tilde{D}(T \cdot \tilde{\tau}_i))_{i \in \mathbb{N}} \tag{4.7}
\]
in \((\mathbb{R}^+)\mathbb{N}^\mathbb{N}\). Again by the joint convergence of the properly normalized and scaled sequential partial sums to the process \((\tilde{D}(\cdot), \tilde{A}(\cdot))\), the independence \((4.4)\) and Theorem \([2]\) convergence of the points

\[
\left( b_n^{-1} T_{d_{[nT]} - 1 + 1} A_{n} X_{[nT] - 1 + 1; [nT]} \right)_{i \in \mathbb{N}} \xrightarrow{\mathcal{D}} (\tilde{D}(T \cdot \tilde{\tau}_i), \tilde{\eta}^{-}(T^{-1} \tilde{\Gamma}_i) \tilde{d}_i, \tilde{V}_i \tilde{d}_i)_{i \in \mathbb{N}}
\]
holds. Since the mapping \(x \mapsto \varepsilon_x\) is continuous, the convergence of points yields the convergence of point processes

\[
\left( \varepsilon \left( b_n^{-1} T_{d_{[nT]} - 1 + 1} A_{n} X_{[nT] - 1 + 1; [nT]} \right) \right)_{i \in \mathbb{N}} \xrightarrow{\mathcal{D}} \left( \varepsilon \left( \tilde{D}(T \cdot \tilde{\tau}_i), \tilde{\eta}^{-}(T^{-1} \tilde{\Gamma}_i) \tilde{d}_i, \tilde{V}_i \tilde{d}_i \right) \right)_{i \in \mathbb{N}}
\]
in \(M_p((0, \infty) \times [-\infty, \infty]^d \setminus \{0\})\). Now summation and an easy application of theorem \(4.2\) in \([4]\) yields

\[
\sum_{k=1}^{\lfloor nT \rfloor} \varepsilon \left( \tilde{d}_{[nT] - k + 1} A_{n} X_{[nT] - k + 1; [nT]} \right) \xrightarrow{\mathcal{D}} \sum_{k \in \mathbb{N}} \varepsilon \left( \tilde{D}(T \cdot \tilde{\tau}_k), \tilde{\eta}^{-}(T^{-1} \tilde{\Gamma}_k) \tilde{d}_k, \tilde{V}_k \tilde{d}_k \right)
\]
in \(M_p((0, \infty) \times [-\infty, \infty]^d \setminus \{0\})\). Now we need to reverse the order of the points again. We introduce

\[
r_k(\omega) := 1 + \# \left\{ i \in \mathbb{N} : \tilde{\eta}^{-}(T^{-1} \tilde{\Gamma}_i(\omega), \tilde{V}_i(\omega)) > \tilde{\eta}^{-}(T^{-1} \tilde{\Gamma}_k(\omega), \tilde{V}_k(\omega)) \right\}
\]
as inverse of \(\tilde{d}_k\). Note that \(r_k\) is well-defined for all \(k \in \mathbb{N}\) by Remark \(3\). Moreover, since \((\tilde{\tau}_n)_{n \in \mathbb{N}}\) and \((r_n)_{n \in \mathbb{N}}\) are independent, an easy application of the desintegration formula shows, that \((\tilde{\tau}_n)_{n \in \mathbb{N}}\) is also i.i.d. and \(\mathcal{U}(0, 1)\) distributed. Since the convergence in \((2.3)\) towards \((2.3)\) can be proven with the same technique, the sequences
\((\tau_n)_{n \in \mathbb{N}}, (\Gamma_n)_{n \in \mathbb{N}}\) and \((\mathbf{V}_n)_{n \in \mathbb{N}}\) are also independent. So we define \(\tau_n := \tau_{rn}\) for all \(n \in \mathbb{N}\) and obtain

\[
\sum_{k=1}^{\lfloor nT \rfloor} \varepsilon(b_n^{-1}T_k\mathbf{A}_n\mathbf{X}_k) = \sum_{k=1}^{\lfloor nT \rfloor} \varepsilon(b_n^{-1}T_{d[nT]-k+1:|nT|}\mathbf{A}_n\mathbf{X}_{|nT|-k+1:|nT|})
\]

\[
\xrightarrow{D} \sum_{k \in \mathbb{N}} \varepsilon(D(T^{-1}\tau_i)\tilde{\eta}^{-1}(T^{-1}\Gamma_k, \mathbf{V}_k)\mathbf{V}_k).
\]

Hence we have proven (4.1). In order to prove (4.2) the limit distribution of

\[
b_n^{-1}T_{d[nT]-i+1-1} = b_n^{-1} \sum_{k=1}^{\lfloor nT \rfloor} J_k = b_n^{-1} \sum_{k=1}^{\lfloor nT \rfloor} J_k
\]

has to be analyzed. Denoting \(G([0, \infty), \mathbb{R}^d)\) the space of all left-continuous functions with right-hand limits from \([0, \infty)\) to \(\mathbb{R}^d\) one easily proves that the mapping \(\mathcal{T} : D[0, \infty) \to G[0, \infty), x(t) \mapsto x(t^-)\) is Lipschitz-continuous with Lipschitz-constant one and hence continuous. Since

\[
\mathcal{T} \left( \sum_{k=1}^{\lfloor nt \rfloor} J_k \right) = \sum_{k=1}^{\lfloor nt \rfloor} J_k
\]

holds, the continuity of \(\mathcal{T}\) suggests, that the convergence

\[
\left( b_n^{-1} \sum_{k=1}^{\lfloor nT \rfloor} J_k \right) \xrightarrow{i \in \mathbb{N}} (D(T \cdot \tau_i -))_{i \in \mathbb{N}}
\]

holds. But since (4.7) could not be proven with the continuous mapping theorem, we use the above arguments again to obtain (4.8). We note that

\[
b_n^{-1} \sum_{l=1}^{\lfloor nT \rfloor} J_l = \sum_{l=1}^{\lfloor nT \rfloor} b_n^{-1} J_{[nT]-l+1:[nt]} \mathbf{1}_{d[nT]-l+1 \leq d[nT]-i+1-1}
\]

\[
\sum_{l=1}^{\lfloor nT \rfloor} b_n^{-1} J_{[nT]-l+1:[nt]} \mathbf{1}_{n^{-1}d[nT]-l+1 < n^{-1}d[nT]-i+1} \xrightarrow{D} \sum_{l \in \mathbb{N}} \tilde{\eta}^{-1}(T^{-1}\Gamma_l) \mathbf{1}_{T: \tau_i < T\tau_i}
\]

holds. Again we identify the limit on the right-hand side in (4.9) as a series representation of \(D(T \cdot \tau_i -)\). As already stated this yields the desired result (4.2), which completes the proof. \(\square\)
5. Scaling limits of coupled CTRWs

In this section we are now able to identify the scaling limits of coupled CTRWs using the limit theorems for their associated point processes stated in Lemma 4. We need to introduce the set

\[ S := \{ T \in \mathbb{R}^+ : \mathbb{P}(D(T \cdot \tau_i) = T) = 0 \text{ for all } i \in \mathbb{N} \} \]

for technical reasons. Due to selfsimilarity, one can easily show that the equality

\[ D(xt \cdot \tau_i) \overset{D}{=} x^{1/\alpha} D(t \cdot \tau_i) \]

holds for all \( i \in \mathbb{N} \) and \( x, t \in \mathbb{R}^+ \). Hence the set \( S \) is dense in \( \mathbb{R}^+ \).

**Theorem 5.** Let \( E(t) := \inf\{ x > 0 : D(x) > t \} \) denote the hitting-time process associated with \( D(\cdot) \). Then convergence of the backward- and forward-coupled CTRW

\[
\sum_{k=1}^{N_{tbn}} (A_nX_k - \mathbb{E}(A_nX_k1_{\|A_nX_k\| \leq \tau}))
\]

(5.1)

\[
\overset{D}{\lim}_{\varepsilon \downarrow 0} \left( \sum_{D(T \cdot \tau_k) \leq t} (\tilde{\eta}^{-}(T^{-1}T_k, V_k) V_k1_{\tilde{\eta}^{-}(T^{-1}T_k, V_k) \geq \varepsilon}) - E(t) \int_{\varepsilon \leq \|x\| \leq \tau} x \, d\eta(x) \right)
\]

and

\[
\sum_{k=1}^{N_{tbn}+1} (A_nX_k - \mathbb{E}(A_nX_k1_{\|A_nX_k\| \leq \tau}))
\]

(5.2)

\[
\overset{D}{\lim}_{\varepsilon \downarrow 0} \left( \sum_{D(T \cdot \tau_k) \leq t} (\tilde{\eta}^{-}(T^{-1}T_k, V_k) V_k1_{\tilde{\eta}^{-}(T^{-1}T_k, V_k) \geq \varepsilon}) - E(t) \int_{\varepsilon \leq \|x\| \leq \tau} x \, d\eta(x) \right)
\]

holds in \( D([0, T], \mathbb{R}^d) \) for every \( T \in S \) and every \( \tau > 0 \) such that \( \eta(S_d^{\tau-1}) = 0 \).

**Proof.** Choose \( T \in S \) arbitrary. Similar to (3.6) we define another a.s. continuous restriction functional

\[ \tilde{\pi}' : M_p([0, \infty) \times [-\infty, \infty]^d \setminus \{0\}) \to M_p([0, \infty) \times [-\infty, \infty]^d \setminus K_d^d), \]

\[ \tilde{\pi}'(m) := m_{[0,\infty) \times [-\infty,\infty]^d \setminus K_d^d} \]
for $\varepsilon > 0$ such that $\eta(S_{\varepsilon}^{d-1}) = 0$. Moreover we define the summation functional

$$
\chi : M_p([0, \infty) \times [-\infty, \infty]^d \setminus \mathbb{R}_x^d) \rightarrow D([0, T], \mathbb{R}^d)
$$

$$
\chi \left( \sum_{k \in \mathbb{N}} \varepsilon(t_k, x_k) \right)(t) = \left( \sum_{t_k \leq t} x_k \right)_{t \in [0, T]}
$$

which is a.s. continuous in the point

$$
\sum_{k \in \mathbb{N}} \varepsilon(D(T \cdot \tau_k), \tilde{\eta}^\tau(T^{-1} \Gamma_k, V_k) V_k 1_{\tilde{\eta}^\tau(T^{-1} \Gamma_k, V_k) > \varepsilon})
$$

for every $T \in S$. A proof of the continuity of $\chi$ is given in [23, Sec. 7.2.3] for the case $d = 1$ and can easily be modified to hold for $d \geq 1$. So we apply the a.s. continuous mapping $\chi \circ \pi$ to the associated point processes in Lemma 4. Considering the equality

$$
\{ T_n \leq t \} = \{ N_t \geq n \} \text{ we receive }
$$

$$
\chi \circ \pi \left( \sum_{k=1}^{[nT]} \varepsilon(b_n^{-1} T_k, A_n X_k) \right)(t) = \left( \sum_{b_n^{-1} T_k \leq t} A_n X_k 1_{\|A_n X_k\| \geq \varepsilon} \right)_{t \in [0, T]}
$$

$$
= \left( \sum_{k=1}^{N_{tb_n}} A_n X_k 1_{\|A_n X_k\| \geq \varepsilon} \right)_{t \in [0, T]} \xrightarrow{D} \chi \circ \pi \left( \sum_{k \in \mathbb{N}} \varepsilon(D(T \cdot \tau_k), \tilde{\eta}^\tau(T^{-1} \Gamma_k, V_k) V_k 1_{\tilde{\eta}^\tau(T^{-1} \Gamma_k, V_k) \geq \varepsilon}) \right)_{t \in [0, T]}
$$

(5.3)

To study the centering constants we use the convergence

$$
\left( \frac{N_{tb_n}}{n} \right)_{t \geq 0} \xrightarrow{D} (E(t))_{t \geq 0}
$$

in $D([0, \infty), \mathbb{R}^+)$, proved in Corollary 3.4 of [21]. Considering (3.5) this yields

$$
\left( \sum_{k=1}^{N_{tb_n}} \mathbb{E} \left( A_n X_k 1_{[\varepsilon, \tau]}(\|A_n X_k\|) \right) \right)_{t \geq 0} = \left( \frac{N_{tb_n}}{n} \int_{\varepsilon \leq \|x\| \leq \tau} x \, nd\mathbb{P}A_n X_k^1(x) \right)_{t \geq 0}
$$

$$
\xrightarrow{D} \left( E(t) \int_{\varepsilon \leq \|x\| \leq \tau} x \, d\eta(x) \right)_{t \geq 0}
$$
in \( D([0, \infty), \mathbb{R}^+) \). Since the process \( E(\cdot) \) has a.s. continuous sample paths, Theorem 4.1 in [30] allows us to put this and (5.3) together. We obtain

\[
\mathbb{P} \left( \sum_{k=1}^{N_{thn}} (A_n \mathbf{X}_k 1_{\|A_n \mathbf{x}_k\| \geq \varepsilon} - \mathbb{E} (A_n \mathbf{X}_k 1_{\varepsilon \leq \|A_n \mathbf{x}_k\| \leq \tau})) \right)_{t \in [0, T]} \to \mathbb{P} \left( \sum_{D(T-T_k) \leq t} \left( \tilde{n}^{-}(\Gamma_k, \mathbf{V}_k) \mathbf{V}_k 1_{\tilde{n}^{-}(T^{-1}\Gamma_k, \mathbf{V}_k) \geq \varepsilon} \right) - \mathbb{E}(t) \int_{\varepsilon \leq \|x\| \leq \tau} x \, d\eta(x) \right)_{t \in [0, T]} \in D([0, T], \mathbb{R}^d). \]

Taking limits as \( \varepsilon \downarrow 0 \) this yields the desired result (5.1). By Theorem 4.2 of [4] it remains to show

\[
\lim_{\varepsilon \downarrow 0} \limsup_{n \to \infty} \mathbb{P} \left( \sup_{0 \leq t \leq T} \left\| \sum_{k=1}^{N_{thn}} (A_n \mathbf{X}_k 1_{\|A_n \mathbf{x}_k\| \geq \varepsilon} - \mathbb{E} (A_n \mathbf{X}_k 1_{\varepsilon \leq \|A_n \mathbf{x}_k\| \leq \tau})) \right\| \geq \delta \right) = 0
\]

for all \( \delta > 0 \). Using a version of the Kolmogorov-inequality for integrable stopping times given in the Appendix and the norm-inequality \( \| \cdot \| \leq \| \cdot \|_1 \) we obtain

\[
\mathbb{P} \left( \sup_{0 \leq t \leq T} \left\| \sum_{k=1}^{N_{thn}} (A_n \mathbf{X}_k 1_{\|A_n \mathbf{x}_k\| \geq \varepsilon} - \mathbb{E} (A_n \mathbf{X}_k 1_{\varepsilon \leq \|A_n \mathbf{x}_k\| \leq \tau})) \right\| \geq \delta \right) \leq \mathbb{P} \left( \max_{1 \leq j \leq N_{thn}} \left\| \sum_{k=1}^{j} A_n \mathbf{X}_k 1_{\|A_n \mathbf{x}_k\| < \varepsilon} - \mathbb{E} (A_n \mathbf{X}_k 1_{\|A_n \mathbf{x}_k\| < \varepsilon}) \right\|_{1} \geq \delta \right) \leq \mathbb{P} \left( \max_{1 \leq j \leq N_{thn}} \left| \sum_{k=1}^{j} (A_n \mathbf{X}_k)^{(i)} 1_{\|A_n \mathbf{x}_k\| < \varepsilon} - \mathbb{E} ((A_n \mathbf{X}_k)^{(i)} 1_{\|A_n \mathbf{x}_k\| < \varepsilon}) \right| \geq \frac{\delta}{d} \right) \leq \sum_{i=1}^{d} \mathbb{P} \left( \max_{1 \leq j \leq N_{thn}+1} \left| \sum_{k=1}^{j} (A_n \mathbf{X}_k)^{(i)} 1_{\|A_n \mathbf{x}_k\| < \varepsilon} - \mathbb{E} ((A_n \mathbf{X}_k)^{(i)} 1_{\|A_n \mathbf{x}_k\| < \varepsilon}) \right| \geq \frac{\delta}{d} \right)
\]
\[
\leq \left( \frac{\delta}{d} \right)^{-2} \mathbb{E}(N_{Tb_n} + 1) \sum_{i=1}^{d} \text{Var} \left( (A_nX_1)^{(i)} 1_{\|A_nX_1\| < \varepsilon} \right)
\]

\[
\leq \left( \frac{\delta}{d} \right)^{-2} \mathbb{E}(N_{Tb_n} + 1) \sum_{i=1}^{d} \mathbb{E} \left( (A_nX_1)^{(i)} 1_{\|A_nX_1\| < \varepsilon} \right)^2 \right),
\]

where \( x^{(n)} \) denotes the \( n \)-th coordinate of the vector \( x \). Note that we have to take \( N_{Tb_n} + 1 \) since \( N_{Tb_n} \) does not fulfill the conditions of Lemma 7. Now Theorem 9 in [19] states that \( \mathbb{E}(N_t + 1) \) can asymptotically be expressed by the integrated tail of the distribution function of \( J_1 \)

\[
\mathbb{E}(N_t + 1) \Gamma(2 - \alpha) \Gamma(1 + \alpha) \sim \frac{t}{\int_0^t (1 - F_{J_1}(x)) dx},
\]

where \( \Gamma \) denotes the gamma-function. By Karamata’s Theorem the limit behavior of the function \( t \mapsto \int_0^t (1 - F_{J_1}(s)) ds \) can be expressed by \( F_{J_1} \)

\[
\int_0^t (1 - F_{J_1}(s)) ds \sim \frac{t(1 - F_{J_1}(t))}{1 - \alpha}.
\]

Putting this together we obtain

\[
\mathbb{E}(N_t + 1) \sim (1 - F_{J_1}(t))^{-1} \cdot \Gamma(1 - \alpha)^{-1} \cdot \Gamma(1 + \alpha)^{-1}.
\]

Defining \( C := (\Gamma(1 - \alpha) \Gamma(1 + \alpha) n^\alpha((T, \infty)))^{-1} \) for abbreviation, the inequality \( |x^{(n)}| \leq \|x\| \) yields

\[
\lim_{\varepsilon \downarrow 0} \lim_{n \rightarrow \infty} \left( \frac{\delta}{d} \right)^{-2} \lim_{\varepsilon \downarrow 0} \lim_{n \rightarrow \infty} \mathbb{E}(N_{Tb_n} + 1) \sum_{i=1}^{d} \mathbb{E} \left( (A_nX_1)^{(i)} 1_{\|A_nX_1\| < \varepsilon} \right)^2 \right)
\]

\[
\leq \lim_{\varepsilon \downarrow 0} \lim_{n \rightarrow \infty} \left( \frac{\delta}{d} \right)^{-2} C \cdot \sum_{i=1}^{d} \lim_{\varepsilon \downarrow 0} \lim_{n \rightarrow \infty} n \mathbb{E} \left( (A_nX_1)^{(i)} 1_{\|A_nX_1\| < \varepsilon} \right)^2.
\]

So it remains to show

\[
\lim_{\varepsilon \downarrow 0} \lim_{n \rightarrow \infty} n \mathbb{E} \left( (A_nX_1)^{(i)} 1_{\|A_nX_1\| < \varepsilon} \right)^2 = 0
\]

for all \( 1 \leq i \leq d \). Defining the tail and truncated second moment of \( X_1 \) in direction \( v \) by

\[
V(r, v) := \mathbb{P}(\|X_1, v\| > r), \quad U(r, v) := \mathbb{E}(\|X_1, v\|^2 1_{\|X_1, v\| < r})
\]

for every \( v \in S^{d-1} \) and every \( r > 0 \) an easy calculation shows that

\[
\mathbb{E} \left( (A_nX_1)^{(i)} 1_{\|A_nX_1\| < \varepsilon} \right)^2 = r_n^2 U(r_n^{-1} \varepsilon, v_n)
\]
holds. Here $r_n > 0$ and $v_n \in S^{d-1}$ are taken such that $A_n^* e_i = r_n v_n$ holds for every $n \in \mathbb{N}$, where $A^*$ is the adjoint of $A$ and $e_1, \ldots, e_d$ denotes the standard basis of $\mathbb{R}^d$.

With this notation we have to analyse

$$n \cdot r_n^2 U (r_n^{-1} \varepsilon, v_n) = \varepsilon^2 \frac{U(r_n^{-1} \varepsilon, v_n)}{\varepsilon^2 r_n^2 V(r_n^{-1} \varepsilon, v_n)} \cdot \frac{V(r_n^{-1} \varepsilon, v_n)}{V(\varepsilon^{-1}(r_n^{-1} \varepsilon), v_n)} \cdot n \cdot V(r_n^{-1}, v_n).$$

One easily verifies that the third factor in (5.4) is bounded by $\eta(\{x \in \mathbb{R}^d : |x(i)| > 1\})$. Since the real parts $a_1 \leq \ldots \leq a_d$ of all eigenvalues of the operator $E$ are greater than $1/2$ we can find $\bar{\varepsilon} > 0$ such that $2 - \bar{\varepsilon} - a_1^{-1} > 0$ holds. So an application of Theorem 6.3.4 of [20] yields the existence of a constant $C_1$, such that for $n \in \mathbb{N}$ large enough the second factor in (5.4) is bounded by $C_1 \varepsilon^{2 - \bar{\varepsilon} - a_1^{-1}}$. Finally, Corollary 6.3.9 in [20] yields that the first factor in (5.4) is bounded by a constant $C_2$. Putting things together this completes the proof.

The proof of the convergence (5.2) works the same way.

Theorem 5 provides a series representation for the limit distribution. This representation might be useful for simulation purposes. Now it is of interest to identify the scaling limits with the ones stated in [29]. The arguments are based on the following equalities:

$$(5.5) \quad \{D(x) < t\} = \{x < E(t)\}, \quad \{D(x) \leq t\} = \{x \leq E(t)\}.$$

The left-hand side of (5.5) is already proven in (3.2) of [21]. For the proof of the right-hand side assume $D(x) \leq t$ holds. So we have $D(y) \leq t$ for all $y < x$. Hence $x \leq E(t)$. Otherwise if $D(x) > t$ holds, there exists an $\varepsilon > 0$ such that $D(y) > t$ holds for all $y \geq x - \varepsilon$. Hence $E(t) \leq x - \varepsilon < x$.

**Corollary 6.** Let $A(t-)^+$ denote the right-continuous version of the process $A(t-)$, i.e. $A(t-)^+$ is an element of $D[0, \infty)$. Then the convergence

$$(5.6) \quad \sum_{k=1}^{N_{tb_n}} (A_n X_k - \mathbb{E}(A_n X_k 1_{\|A_n X_k\| \leq \tau})) \xrightarrow{D} A(E(t)-)^+$$

and

$$(5.7) \quad \sum_{k=1}^{N_{tb_n}+1} (A_n X_k - \mathbb{E}(A_n X_k 1_{\|A_n X_k\| \leq \tau})) \xrightarrow{D} A(E(t))$$

holds in $D([0, \infty), \mathbb{R}^d)$ for any $\tau > 0$ such that $\eta(S_{d}^{-1}) = 0$. 

Proof. The convergence (5.7) can easily be verified applying the right-hand side of (5.5) to the series representation (2.3). To prove (5.6) we apply the left-hand side of (5.5) to (5.1) and obtain
\[
\sum_{k=1}^{N_{tn}} (A_n X_k - \mathbb{E}(A_n X_k)) \xrightarrow{D} \lim_{\varepsilon \downarrow 0} \left( \sum_{D(T \cdot \tau_k) \leq t} (\tilde{\eta}^{-}(T^{-1} \Gamma_k, V_k) V_k 1_{\tilde{\eta}^{-}(T^{-1} \Gamma_k, V_k) \geq \varepsilon}) - E(t) \int_{\varepsilon \leq \|x\| \leq \tau} x \, d\eta(x) \right)
\]
(5.8)
\[
+ \sum_{D(T \cdot \tau_k) = t} (\tilde{\eta}^{-}(T^{-1} \Gamma_k, V_k) V_k 1_{\tilde{\eta}^{-}(T^{-1} \Gamma_k, V_k) \geq \varepsilon}) - E(t) \int_{\varepsilon \leq \|x\| \leq \tau} x \, d\eta(x)
\]
As we already stated
\[
\lim_{\varepsilon \downarrow 0} \left( \sum_{T \cdot \tau_k < E(t)} (\tilde{\eta}^{-}(T^{-1} \Gamma_k, V_k) V_k 1_{\tilde{\eta}^{-}(T^{-1} \Gamma_k, V_k) \geq \varepsilon}) - E(t) \int_{\varepsilon \leq \|x\| \leq \tau} x \, d\eta(x) \right)
\]
is a series representation of \( A(E(t) -) \). The extra summands only have to be considered if a jump occurs at a time \( t \) with \( D(T \cdot \tau_k) = t \). This yields the right-continuity of the limit and we have proven (5.6). \( \square \)

Appendix A. A generalization of Kolmogorov’s inequality

The following generalisation of Kolmogorov’s inequality can be shown by standard techniques. However, we were not able to find a suitable proof in the literature and will only give a sketch of proof.

Lemma 7. Let \((Y_n)_{n \in \mathbb{N}}\) be i.i.d. with \( \mathbb{E}(Y_1) = 0 \) and \( T \) be an \( \mathbb{N}_0 \)-valued integrable stopping time with respect to the filtration \( \mathcal{F}_n := \sigma(Y_1, \ldots, Y_n) \). Then
\[
P \left( \max_{1 \leq k \leq T} \left| \sum_{j=1}^{k} Y_j \right| \geq \delta \right) \leq \delta^{-2} \cdot \mathbb{E}(T) \cdot \text{Var}(Y_1)
\]
holds for all \( \delta > 0 \).
**Proof.** First we restrict our attention to the truncated stopping time $T \land n$. An easy calculation shows that
\[
\mathbb{E} \left( \sum_{k=1}^{T \land n} (S_k - S_{k-1})^2 \right) = \mathbb{E}(S_{T \land n}^2)
\]
holds, where $S_n := \sum_{k=1}^{n} Y_k$ denotes the $n$-th partial sum. Defining
\[
M_0 := 0, \quad M_{k+1} := \begin{cases} S_{k+1}, & \text{if } \max_{1 \leq j \leq k} |S_j| < \delta \\ M_k, & \text{else} \end{cases}
\]
the same calculation shows
\[
\mathbb{E} \left( \sum_{k=1}^{T \land n} (M_k - M_{k-1})^2 \right) = \mathbb{E}(M_{T \land n}^2) - 2 \cdot \mathbb{E} \left( \sum_{k \in \mathbb{N}} (M_k - M_{k-1})S_{k-1} 1_{k \leq T \land n} \right).
\]
Using the definition of $M_n$ one can show
\[
\mathbb{E} \left( \sum_{k=1}^{T \land n} (M_k - M_{k-1})^2 \right) = \mathbb{E}(M_{T \land n}^2).
\]
Now $|M_k - M_{k-1}| \leq |S_k - S_{k-1}|$ holds for all $k \in \mathbb{N}$. Hence the Markov-inequality yields
\[
P \left( \max_{1 \leq k \leq T \land n} |S_k| \geq \delta \right) = P \left( |M_{T \land n}| \geq \delta \right) \leq \delta^{-2} \mathbb{E}(M_{T \land n}^2)
\]
\[
\leq \delta^{-2} \mathbb{E} \left( \sum_{k=1}^{T \land n} (M_k - M_{k-1})^2 \right) \leq \delta^{-2} \mathbb{E} \left( \sum_{k=1}^{T \land n} (S_k - S_{k-1})^2 \right) = \delta^{-2} \mathbb{E}(S_{T \land n}^2).
\]
An application of Wald’s inequality yields the desired result for $T \land n$. The generalisation for the stopping time $T$ follows by the martingale convergence theorem. \qed

**Remark 8.** Let $(Z_n)_{n \in \mathbb{N}}$ be a sequence such that $(Y_n, Z_n)_{n \in \mathbb{N}}$ is i.i.d.. Then Lemma also holds replacing $\mathcal{F}_n$ by $\tilde{\mathcal{F}}_n := \sigma((Y_1, Z_1), \ldots, (Y_n, Z_n))$.

**References**

[1] A. Barczyk, A. Janssen and M. Pauly, The asymptotics of L-statistics for non i.i.d. variables with heavy tails, *Probab. Math. Stat.* 31 (2011), 285–299.
[2] P. Becker-Kern, Random sums of independent random vectors attracted by (semi)-stable hemi-groups, *J. Appl. Anal.* 10 (2004), 83–104.
[3] P. Becker-Kern, M.M. Meerschaert and H.P. Scheffler, Limit theorems for coupled continuous time random walks, *Ann. Probab.* 32 (2004), 730–756. Correction: *Ann. Probab.* 40 (2012), 890–891.
[4] P. Billingsley, *Convergence of Probability Measures*, 2nd Edition, Wiley Series in Probability and Statistics (1999), Wiley, New York.
[5] P. Feigin, M.F. Krantz and S.I. Resnick, Parameter estimation for moving averages with positive innovations, *Ann. Appl. Probab.* 6 (1996), 1157–1190.
[6] T.S. Ferguson and M.J. Klass, A representation of independent increment processes without Gaussian components, *Ann. Math. Stat.* 43 (1972), 1634–1643.
[7] M. Finkelstein and H.G. Tucker, A necessary and sufficient condition for convergence in law of random sums of random variables under nonrandom centering, *Proc. Amer. Math. Soc.* 107 (1989), 1061–1070.
[8] B. Gnedenko, and A.V. Kolmogorov, *Limit Distributions for Sums of Independent Random Variables*, Addison-Wesley (1968), Reading, MA.
[9] M.G. Hahn, W.N. Hudson und J.A. Veeh, Operator stable laws: Series representation and domains of normal attraction, *J. Theor. Probab.* 2 (1989), 3–35.
[10] A. Janßen and D.M. Mason, Non-standard rank tests, *Lect. Notes Stat.* 65 (1990), Springer, New York.
[11] M. Jara and T. Komorowski, Limit theorems for some continuous-time random walks, *Adv. Appl. Probab.* 43 (2011), 782–813.
[12] A. Jurlewicz, P. Kern, M.M. Meerschaert and H.P. Scheffler, Fractional governing equations for coupled random walks, preprint (2011), to appear in *Comp. Math. Appl.*, doi:10.1016/j.camwa.2011.10.010.
[13] O. Kallenberg, *Random Measures*, Akad.-Verl. (1986), Berlin.
[14] J. Klafter, A. Blumen and M.F. Shlesinger, Stochastic pathways to anomalous diffusion, *Phys. Rev. A* 35 (1987), 3081–3085.
[15] M. Kotulski, Asymptotic distributions of the continuous time random walks: A probabilistic approach, *J. Statist. Phys.* 81 (1995), 777–792.
[16] E.L. Lehmann and J.P. Romano, *Testing Statistical Hypotheses*, Springer (2005), New York.
[17] R. LePage, Multidimensional infinitely divisible variables and processes part II, *Lect. Notes Math.* 860 (1981), 297–284.
[18] R. LePage, M. Woodroofe and J. Zinn, Convergence to a stable distribution via order statistics, *Ann. Probab.* 9 (1981), 624–632.
[19] F. Mallor and E. A.M. Omey, *Univariate and Multivariate Weighted Renewal Theory*, Public University of Navarra (2006), Pamplona.
[20] M.M. Meerschaert and H.P. Scheffler, *Limit Distributions for Sums of Independent Random Variables*, Wiley Series in Probability and Statistics (2001), Wiley, New York.
[21] M.M. Meerschaert and H.P. Scheffler, Limit theorems for continuous-time random walks with infinite mean waiting times, *J. Appl. Probab.* 41 (2004), 623–638.
[22] E. Montroll and G.H. Weiss, Random walks on lattices. II, *J. Math. Phys.* 6 (1965), 167–181.
[23] S.I. Resnick, *Heavy-Tail Phenomena*, Springer (2007), New York.
[24] J. Rosiński, Series representations of Lévy processes from the perspective of point processes, In: O.E. Barndorff-Nielsen et al. (eds.) *Lévy Processes - Theory and Applications, Birkhäuser (2001), Boston, pp. 401–415.
[25] E. Scalas, Five years of continuous-time random walks in econophysics, *Proc. WEHIA* 55 (2004), 223–230.
[26] H.P. Scheffler, Series representation for operator semistable laws and domains of normal attraction, *J. Math. Sci.* 92 (1998), 4062–4084.
[27] R. Schumer, M.M. Meerschaert and B. Baeruem, Fractional advection-dispersion equations for modeling transport at the Earth surface, *J. Geophys. Res.* (2009) 114, F00A07.
[28] M. Shlesinger, J. Klafter and Y.M. Wong, Random walks with infinite spatial and temporal moments, *J. Stat. Phys.* 27 (1982), 499–512.
[29] P. Straka and B.I. Henry, Lagging and leading coupled continuous time random walks, renewal
times and their joint limits, Stoch. Process. Appl. 121 (2011), 324–336.
[30] W. Whitt, Some useful functions for functional limit theorems, Math. Oper. Res. 5 (1980), S.
67–85.

Mathematical Institute, Heinrich-Heine-University Düsseldorf, Germany
E-mail address: barczyk@math.uni-duesseldorf.de

Mathematical Institute, Heinrich-Heine-University Düsseldorf, Germany
E-mail address: kern@math.uni-duesseldorf.de