The Sloan Digital Sky Survey Reverberation Mapping Project: Accretion Disk Sizes from Continuum Lags

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Abstract

We present accretion disk structure measurements from continuum lags in the Sloan Digital Sky Survey Reverberation Mapping (SDSS-RM) project. Lags are measured using the JAVELIN software from the first-year SDSS-RM g and i photometry, resulting in well-defined lags for 95 quasars, 33 of which have lag S/N > 2σ. We also estimate lags using the ICCF software and find consistent results, though with larger uncertainties. Accretion disk structure is fit using a Markov chain Monte Carlo approach, parameterizing the measured continuum lags as a function of disk size normalization, wavelength, black hole mass, and luminosity. In contrast with previous observations, our best-fit disk sizes and color profiles are consistent (within 1.5σ) with the Shakura & Sunyaev analytic solution. We also find that more massive quasars have larger accretion disks, similarly consistent with the analytic accretion disk model. The data are inconclusive on a correlation between disk size and continuum luminosity, with results that are consistent with both no correlation and the Shakura & Sunyaev expectation. The continuum lag fits have a large excess dispersion, indicating that our measured lag errors are underestimated and/or our best-fit model may be missing the effects of orientation, spin, and/or radiative efficiency. We demonstrate that fitting disk parameters using only the highest-S/N lag measurements biases best-fit disk sizes to be larger than the disk sizes recovered using a Bayesian approach on the full sample of well-defined lags.

Key words: accretion, accretion disks – galaxies: active – galaxies: nuclei – quasars: general

Supporting material: figure set, machine-readable table

1. Introduction

Quasars are supermassive black holes (SMBHs) that grow by rapid mass accretion. During the accretion phase, quasars glow with total luminosity \( L_{\text{bol}} = \eta M \dot{M} c^2 \), where \( \eta \) is the radiative efficiency, \( \dot{M} = \frac{dM}{dt} \) is the SMBH accretion rate, and \( c \) is the speed of light. The foundational model for black hole accretion disks is the thin-disk model of Shakura & Sunyaev (1973, hereafter SS73). The SS73 disk model is an optically thick, geometrically thin disk model where the local disk emission corresponds to a series of blackbodies at different radii. The inner part of the accretion disk has hotter emission, whereas at the outer edge of the disk, the emission is cooler.

Even though the SS73 model has been widely used, mounting observational and theoretical evidence shows that the SS73 disk model breaks down in several ways. Recent continuum reverberation mapping (RM) observations (Shappee et al. 2014; Fausnaugh et al. 2016, 2017; Jiang et al. 2017a; Mudd et al. 2018) identified discrepancies in the measured disk sizes from what is expected by the SS73 model. This discrepancy is also reported in microlensing observations of quasars (Morgan et al. 2010; Blackburne et al. 2011; Motta et al. 2017).

Both theory and non-RM observations suggest that black hole accretion structure depends on accretion rate in ways that are not entirely predicted by the SS73 thin-disk model. Recent advances in simulations of super-Eddington accretion disks predict dramatically different emission and outflow properties compared to the sub-Eddington SS73 analytic prescription (Jiang et al. 2014, 2017b; McKinney et al. 2014; Sadowski et al. 2014, 2016; see also the analytic “slim” disk model of Abramowicz et al. 1988). Observations of candidate super-Eddington quasars in the X-ray (Desroches et al. 2009) with broad-line kinematics (Du et al. 2015) and spectral energy distribution (SED) fitting (Luo et al. 2015) show similar evidence for slim accretion disks. At low accretion rates, SED observations suggest that accretion occurs in a hot, ionized, optically thin, radiatively inefficient accretion flow mode, although the exact radiative efficiency is degenerate with the
(2011) incorporate temperature fluctuations in Keplerian rotation disks that can produce larger disk sizes; in addition, this would solve the problem of quasar variability that is not well understood in the context of the SS73 model. However, previous studies have not tested disk structure dependency on \( M_{\text{BH}} \) and accretion rate due to current data limited to low-luminosity Seyfert galaxies. There are currently only seven Type 1 Seyfert active galactic nuclei (AGNs) that have both continuum and emission line RM measurements, which together allow for both direct \( M_{\text{BH}} \) and accretion size measurements (Collier et al. 1998; Edelson et al. 2015, 2017; Fausnaugh et al. 2016; McHardy et al. 2018).

We address this problem by performing a comprehensive study of the physics of black hole accretion using direct accretion disk size and structure measurements from the Sloan Digital Sky Survey Reverberation Mapping (SDSS-RM) project (Shen et al. 2015) between the optical \( g \) and \( i \) photometry bands. We connect the observed accretion disk structure with black hole mass and accretion rate using the JAVELIN software, which fits reverberation lags using a damped random walk (DRW) model for the statistical behavior of light-curve variability. In contrast, Starkey et al. (2019, in preparation), which uses a different methodology to similarly measure continuum lags from SDSS-RM quasars. Here we focus on measuring disk size, color profile, and disk dependence on mass and luminosity using the JAVELIN software, which fits reverberation lags using a damped random walk (DRW) model for the statistical behavior of light-curve variability. In contrast, Starkey et al. (2019, in preparation) used the CREME software with models for both the driving light curve and the disk reverberation response, fitting disk size, temperature profile, and orientation. Section 2 describes our sample chosen from the SDSS-RM data set. Section 3 presents our procedure for lag identification, including alias removal, outlier rejection, and lag quality analysis. In Section 4 we discuss the necessary criteria for selecting physical lags corresponding to reverberating light curves. Section 5 describes our use of computed lags to fit a normalization of the accretion disk and link the observed lags to mass and accretion rate correlations. Throughout this work, we adopt a \( \Lambda \)CDM cosmology with \( \Omega_\Lambda = 0.7 \), \( \Omega_\text{M} = 0.3 \), and \( h = 0.7 \).

2. Data

2.1. SDSS-RM Survey

The SDSS-RM project is a pioneering multi-object RM campaign (Shen et al. 2015) that has been simultaneously monitoring a sample of 849 quasars in a single 7 deg\(^2\) field.
since 2014. The project began with SDSS-III (Eisenstein et al. 2011). The selected RM sample is flux-limited to $I_{\text{psf}} = 21.7$, with no additional cuts on variability amplitude or redshift of the quasars, dramatically expanding the parameter space of the spectroscopic, variability, and multiwavelength properties of quasars with RM data (Figure 1 of Shen et al. 2015). The main goal of SDSS-RM is to measure black hole mass and lags for a range of emission lines, as well as improving the established radius–luminosity (R-L) relation (Kaspi et al. 2007; Bentz et al. 2013) that is currently well calibrated for Hβ in a biased sample of nearby $z < 0.3$ quasars. Due to the necessity of continuous observations in this survey, coordinated monitoring by different SDSS-RM photometry sites is essential to monitor quasar light variability. Thus, the SDSS-RM program is supported by ground-based photometry from multiple facilities, including the Canada–France–Hawaii Telescope (CFHT) and Steward Observatory Bok telescope. To date, SDSS-RM has resulted in several studies of the variability and properties of quasar emission lines (Sun et al. 2015, 2018b; Denney et al. 2016a, 2016b; Shen et al. 2016a; Li et al. 2017), broad absorption line variability (Grier et al. 2016), the relationship between black hole growth and host galaxy properties, and broad emission line lags (Matsuoka et al. 2015; Shen et al. 2016b; Grier et al. 2017).

Here we select the 222 quasars in SDSS-RM (see Figures 1 and 2) with $z < 1.13$ previously studied for broad-line RM and black hole mass, $M_{\text{BH}}$, estimates (Grier et al. 2017). Of the 222 quasars, 44 have reliable $M_{\text{BH}}$ estimates from Grier et al. (2017), enabling us to study the accretion disk structure dependence on black hole mass. The selected sample is unique, as it has well-measured black hole masses and is suitable to study accretion disk properties based on continuum lag measurements.

### 2.2. Spectroscopy

We use the Baryon Oscillation Spectroscopic Survey spectrograph (Dawson et al. 2013; Smee et al. 2013) covering wavelengths of 3650–10400 Å with a spectral resolution of $R \sim 2000$ with the spectrograph mounted on the 2.5 m SDSS telescope (Gunn et al. 2006). Our study uses the first year of SDSS-RM spectroscopic observations, obtained during seven dark/gray observing windows in 2014 January–July. Each epoch has a typical depth of $S/N^2 > 20$ (the average extinction-corrected $S/N^2$ per pixel in the g band evaluated at $g_{\text{psf}} = 21.2$; Shen et al. 2015), with a total of 32 spectroscopic epochs separated by a median of 4 days with varying cadence depending on weather conditions and scheduling constraints.

The spectroscopic data processing is initially processed using the standard SDSS pipeline (Bolton et al. 2012) for flat-fielding, 1D extraction, wavelength calibration, and a first pass at sky subtraction and flux calibration. The SDSS-RM data are also processed with a second round of sky subtraction and flux calibration using a custom pipeline that uses position-dependent calibration vectors (see Shen et al. 2015 for details). Finally, a software package called PrepSpec is used to model the spectra and remove any remaining epoch-dependent calibration errors. This step is implemented by fitting a simple model for quasar spectra and considering a wavelength- and time-dependent component to the continuum and a nonvariable
component to the narrow emission line fluxes. See Shen et al. (2016b) for details.

We measure synthetic photometry in the $g$ and $i$ bands by integrating the SED with the SDSS filter response function (Fukugita et al. 1996; Doi et al. 2010) and the flux errors. The synthetic flux error is computed using the quadratic sum of errors in the measured SED, the shape of the response function, and PrepSpec calibration.

Following Grier et al. (2017), we excluded epoch 7 (MJD = 56,713) from the 32 available epochs because it was taken under poor observing conditions, had significantly lower S/N, and was frequently ($\geq1/3$ of the time) a $>1\sigma$ outlier compared to the other epochs. Furthermore, to improve the overall quality of the obtained continuum light curves, a small number of epochs (1%) are rejected as outliers if offset from the median flux by more than five times the normalized median absolute deviation. This is implemented to mostly remove data points where the fibers were incorrectly placed, altering the flux or dropped fibers.

2.3. Photometry

The SDSS-RM is supported by ground-based photometry from the 3.6 m CFHT and the 2.5 m Steward Observatory Bok telescope. Between 2014 January and June, the Bok/90 Prime instrument (Williams et al. 2004) obtained 31 epochs in the $g$ band and 27 epochs in the $i$ band during 60 observing nights in bright time. The CFHT MegaCam (Aune et al. 2003) obtained 26 epochs in $g$ and 20 epochs in $i$.

The photometric light curves are computed using image subtraction as implemented in the ISIS package (Alard 2000). ISIS first creates a reference image using the best seeing exposure, then matches the astrometry of subsequent frames with different point-spread functions (PSFs). This step uses a least-squares fit to find the optimal kernel between the reference image and the target image while accounting for PSF variation in each target image. The target image is then convolved and subtracted from the reference image to produce the light curves. The reference image and image subtraction is performed for each individual telescope, filter, CCD, and field (K. Kinemuchi et al. 2019, in preparation).

2.4. Light-curve Merging

The combined monitoring from the SDSS, Bok, and CFHT telescopes provides a total of 88 epochs of $g$-band photometry and 78 epochs of $i$-band photometry (see Figure 3). The mean fractional variability is 8.4% in the $g$ band and 7.3% in the $i$ band, in both cases calculated as the maximum-likelihood intrinsic variability accounting for the observational uncertainties (following Almaini et al. 2000; Sun et al. 2015). However, combining the three light curves is nontrivial, since each observatory has different seeing conditions and calibration issues for each filter response, telescope throughput, and any other site-dependent calibration. We use the Continuum REprocessing AGN Markov chain Monte Carlo (CREAM) software (Starkey et al. 2016) to intercalibrate the light curves obtained at different sites with the following model:

$$F_e(t) = F + \Delta F \langle X \rangle,$$

where the light-curve shape $X(t)$ is normalized to $\langle X \rangle = 0$ and $\langle X^2 \rangle = 1$ so that $\langle X \rangle$ is the mean and $\Delta F \langle X \rangle$ is the rms flux of the light curve. CREAM uses a power-law prior on the power spectrum of $X(t)$ (see Equations (8)–(10) of Starkey et al. 2017) so that $\langle X \rangle$ by default resembles the observed behavior of AGN light curves (see Grier et al. 2017 for a step-by-step description of the light-curve merging procedure). The fit allows $\langle X \rangle$ and $\Delta F \langle X \rangle$ to be different for the data from each site while applying the same $\langle X \rangle$ to all sites. The site-to-site differences in $\langle X \rangle$ and $\Delta F \langle X \rangle$ then allow the data from each site to be scaled and shifted and thereby effectively merged into a single light-curve data set with a common photometric calibration. This was done independently for the $i$ and $g$ photometry, thus defining a (slightly) different $\langle X \rangle$ for each band.

3. Continuum RM Analysis

The SDSS-RM light curves are irregularly sampled due to weather conditions and constraints on telescope allotted time; thus, the RM analysis requires interpolation between epochs. We use two approaches to interpolate and measure lags and uncertainties from the merged light curves.

3.1. ICCF

Our first RM analysis methodology is the interpolated cross-correlation function (ICCF; Gaskell & Sparke 1986; Gaskell & Peterson 1987; White & Peterson 1994; Peterson 2004), where observations from different epochs are linearly interpolated to create an evenly sampled grid and calculate the Pearson coefficient $r$ between the two mean-subtracted light curves $S_1(t)$ and $S_2(t)$. The first light curve is then shifted by a time lag $\tau$, and $r$ is remeasured. This step is repeated across the range of allowed $\tau$, thus constructing the CCF. The same procedure is repeated by shifting the other light curve by all $\tau$ values, and the final correlation function is averaged between the two. Determining well-measured lags using the ICCF method is challenging, considering the correlated errors associated with the light-curve interpolation. We estimate errors on the ICCF lags using Monte Carlo simulations. The mean and median of these errors are 3 and 2 days for RM 267, respectively.
Carlo (MC) iterations for flux resampling and random subset selection (Peterson 2004), implemented using the publicly available PyCCF software (Sun et al. 2018a). The flux in each point is resampled by a Gaussian distribution determined by its uncertainty, a random subset of epochs is chosen (with replacement), and the lag is recomputed. Repeated MC is used to obtain cross-correlation peak distribution. The centroid of the CCF is restricted to the region where the CCF is above an 80% fraction of the peak; experimentation reveals that using the centroid of the CCF rather than the CCF peak results in less biased lags and yields higher precision in virial masses (Peterson 2004); we thus choose to work with cross-correlation centroid distribution (CCCD).

We adopt a delay grid spanning ±100 days with a spacing of half the mean of the minimum separation between observed epochs. This search baseline is roughly half the total 180 day range of the SDSS-RM observations and effectively prevents matching nonoverlapping features between the light curves. We perform 5000 MC iterations over the range of allowed $\tau$ per light curve, returning the CCCD for the lag centroid $\tau_{\text{cent}}$ and the cross-correlation Pearson coefficient $r$ at each time delay within the range. Each of the ICCF MC realizations is tested for correlation coefficient and significance of the lag and returns a “failed peak” if significance criteria are not met (i.e., the CCF peak is found to be on the upper or lower limit of the delay grid or the correlation coefficient is less than 0.2 for data points within the centroid). Out of the unique sample of 222 RM objects, RM 173 showed the most failed peak detection with only 37 successfully detected peaks out of 5000 MC realizations. We therefore exclude this quasar, as its CCCD is not statistically significant (we will shortly see that JAVELIN is also unable to obtain the continuum model for RM 173). In the rest of our sample, ~30% of the objects have all 5000 successful MC realizations, and on average, each object has an ~85% success rate.

3.2. JAVELIN

We also compute lags using the JAVELIN software (Zu et al. 2011). JAVELIN assumes a DRW model to predict the light curves at unmeasured times. Observations confirm that the DRW model is a reasonable first-order description of quasar light-curve variability on timescales of $\gtrsim$1 day, with the variability amplitude and damping timescale$^{16}$ dependent on quasar luminosity (Kelly et al. 2009; Kozłowski et al. 2010; MacLeod et al. 2010; Sun et al. 2018c). The DRW in the continuum is first modeled by two priors to compute the continuum light-curve variability with the assumption of covariance between times $t_i$ and $t_j$:

$$<S_i(t_i)S_j(t_j)> = \sigma^2 (1 - e^{-|t_i-t_j|/\tau_d}) \, . \tag{3}$$

Here $\tau_d$ is the damping timescale. This variability model can be approximated as a double power law with a short timescale ($\Delta t < \tau_d$) rms of $\sigma \sqrt{2} \Delta t / \tau_d$ power spectrum power law of $\alpha = -2$ and a long timescale rms of $\sigma (\alpha = 0)$.

JAVELIN models the reverberation response $\Psi(\tau)$ as a top-hat function centered at $\tau$ with full width $\Delta \tau$. The reverberating light curve is then the “lagged” version of the driving light curve smoothed and scaled by the parameters of the top-hat function.

JAVELIN uses a two-step Markov chain MC (MCMC) simulation (Zu et al. 2011). The first step analyzes the driving light curve by itself and obtains uncertainties and posterior distributions for the DRW parameters $\tau_d$ and $\sigma$. The second MCMC analysis determines the best-fit transfer function centroid $\tau$ and $\Delta \tau$ based on the posterior distribution from the isolated continuum in the first MCMC, where each DRW parameter is the median value with the Gaussian width chosen to match the upper and lower $1\sigma$ confidence regions. This approach results in three new posteriors: the mean lag $\tau = (\tau_i + \tau_f)/2$, the width of the top-hat $\Delta \tau = \tau_f - \tau_i$, and a scaling coefficient $A$. The second MCMC process also updates the posterior distribution for the DRW parameters $\tau_d$ and $\sigma$. JAVELIN is able to allow for all of the parameters of the DRW model and transfer function to vary in the MCMC; however, we fix the damping timescale $\tau_d = 200$ days. The assumed damping timescale does not affect the model light curves so long as it is longer than our 180 day monitoring duration. We similarly fix the transfer function to have a width of $\Delta \tau = 0.5$ day, after testing values between 0.25 and 1 day and finding no significant differences in the measured lags. The top-hat function used by JAVELIN is a simplification of the more complicated transfer function likely to describe accretion disk reprocessing (see Starkey et al. 2016), but it is a reasonable approximation so long as the disk response is short compared to the lag, and it has been the common assumption of previous work to which we make comparisons. The uncertainty of the DRW parameters is obtained based on the statistical confidence limits from the posterior distribution. JAVELIN fails to compute the continuum model for RM 173 just as the ICCF failed, and it also fails to compute the continuum model for RM 187 and RM 846. In the end, we have 219 quasars that have computed JAVELIN lags.

We demonstrate the continuum lag analysis results in Figure 4 for RM 267 for the g- and i-band continuum model using JAVELIN and ICCF. Similar figures for our full sample are provided as a figure set.

3.3. Lag Identification Method

Identifying a well-measured lag from the methods described in Sections 3.1 and 3.2 requires additional checks to eliminate cases that appear to be unreliable or ambiguous. Additionally, in many cases, the CCCDs obtained from our methods have multiple peaks that correspond to aliases in the lags due to semi-repeating features in the light curves. Also, it is not always clear if the initial reported lag corresponds to genuine reverberation. We devise a set of criteria to identify unambiguous lags likely to correspond to real reverberation while rejecting less reliable lags.

3.3.1. Aliasing Removal

As mentioned above, many of our quasars have CCCDs with multiple peaks corresponding to competing alternatives for the CCF lag. Some of these peaks occur at the bounds of the time window (±100 days) and are caused by numerical issues.

We assume a prior that lags are most likely to be detected when the two light curves have maximal overlap. Conversely, if shifting epochs by a time delay results in zero overlapping data points between common epochs, then the probability of
finding a lag will be zero. We adopt the same weighting and alias removal scheme as in Grier et al. (2017). The weight is defined as $P(\tau) = \left(\frac{N(\tau)/N_0}{2}\right)^2$, with $N(\tau)$ corresponding to the number of overlapping epochs between the $g$ and $i$ light curves shifted by lag $\tau$ and $N_0$ corresponding to the maximum number of overlapping epochs from the $g$ and $i$ light curves at zero time delay $\tau = 0$.

Our general framework for finding lags is based on the JAVELIN posterior distribution as the CCCD. The CCCD is weighted by $P(\tau)$ to avoid alias lag solutions and smoothed using a Gaussian filter with a width of 5 days. The smoothing is used to identify peaks in the weighted CCCD, as well as the local minima around each peak. The weighted, smoothed CCCD may contain multiple peaks, with a high-significance peak accompanied by multiple low-significance peaks. We compute the area between consecutive local minima and identify the local minima that contain the peak with the most area and adopt the lag as the median of the unsmoothed CCCD within the identified local minima. Furthermore, this technique is helpful in identifying more plausible lags for those CCCDs that show peaks on either end of the lag interval.

The lag uncertainty is computed as the mean absolute deviation relative to the median, computed between the local minima on either side of the peak.

3.3.2. BLR Impact on Continuum Light Curves

The $g$ and $i$ photometric bands in our light curves may include substantial flux from broad emission lines in addition to the continuum emission. Considering that BLR lags typically have longer timescales and show smaller-amplitude variability compared to continuum lags (MacLeod et al. 2012), BLR contamination may potentially affect the observed time lag derived from the continuum. We consider emission lines that could fall in the range of SDSS filters depending on the redshift of our quasar sample: C IV, C III, Mg II, H$\beta$, and H$\alpha$ at, respectively, 1550, 1909, 2799, 4861, and 6563 Å in the rest frame. We determined the broad-line contribution, $f_{\text{BLR}}$, in each as the ratio of emission line equivalent width (Shen et al. 2019) to SDSS filter effective width (Fukugita et al. 1996). The contamination result for all of the objects in our sample is illustrated in Figure 5.

Figure 4. Continuum $g$- (blue) and $i$-band (red) light curves and errors for quasar RM 267 computed with JAVELIN. For clarity, black points indicate the averages of data taken within a single night, although all lag analyses were performed on the individual, nonaveraged observations displayed by small gray points. The best-fit JAVELIN DRW models are shown by the shaded lines in each panel. Bottom left: cross-correlation coefficient computed at each lag with its maximum identified by a red horizontal line. Bottom middle: lag probability distribution computed by ICCF, with the local minima of the primary peak indicated by gray shading and the identified lag and $\pm 1\sigma$ error indicated by the green dotted line and shading. Bottom right: lag probability distribution computed by JAVELIN. The main lag and its $\pm 1\sigma$ error are represented by the red dotted line and shading. In both plots, the Gaussian-smoothed curve represents the smoothed peak with a 5 day standard deviation.

(The complete figure set (219 images) is available.)

Figure 5. Contamination by different broad emission lines in the $g$-band (top) and $i$-band (bottom) photometry of our 222 quasars obtained from Shen et al. (2019). Broad-line contamination, $f_{\text{BLR}}$, is calculated as EW(line)/FWHM(band). We require $<12.5\%$ broad-line contamination for a well-defined photometric accretion disk lag. As shown in the bottom panel of Figure 6, few of the quasars have more than 12.5\% maximum contamination in the $g$ and $i$ bands.
Table 1
Well-defined Quasar Sample Information

| RMID | R.A. (deg) | Decl. (deg) | z  | \( M_{BH}\) (\( M_\odot \)) | \( \log \lambda L_{\lambda,3300} \) (erg s\(^{-1}\)) | \( \tau_{ICCF} \) (days) | \( \tau_{JAVELIN} \) (days) | S/N(\( \tau_{JAVELIN} \)) |
|------|------------|-------------|----|-----------------|-----------------|-----------------|-----------------|-----------------|
| 016  | 214.0290   | 53.1583     | 0.848 | 9.07 \(^\pm\) 0.22 | 44.85 | \(-3.76 \pm 0.74\) | \(-4.01 \pm 0.71\) | \(-3.07\) |
| 017  | 213.3511   | 53.0908     | 0.456 | 8.92 \(^\pm\) 0.24 | 44.16 | \(2.93 \pm 1.24\) | \(5.50 \pm 1.62\) | \(3.30\) |
| 029  | 213.2946   | 52.9640     | 0.816 | 7.72             | 44.12 | \(-1.00 \pm 1.70\) | \(0.33 \pm 1.29\) | \(0.26\) |
| 061  | 214.0000   | 52.7378     | 0.983 | 8.18             | 44.44 | \(-2.54 \pm 0.76\) | \(10.01 \pm 2.60\) | \(3.86\) |
| 062  | 213.5737   | 53.4697     | 0.808 | 8.64\*           | 44.25 | \(-1.18 \pm 2.85\) | \(0.46 \pm 1.67\) | \(0.27\) |
| 078  | 212.9757   | 53.1887     | 0.581 | 8.88\*           | 44.57 | \(-0.11 \pm 1.98\) | \(3.57 \pm 3.79\) | \(0.94\) |
| 088  | 212.9657   | 52.8956     | 0.516 | 8.51             | 44.25 | \(-0.47 \pm 1.72\) | \(-0.25 \pm 2.84\) | \(-0.34\) |
| 101  | 213.0592   | 53.4296     | 0.458 | 7.26\(^\pm\) 0.17 | 44.64 | \(-1.28 \pm 1.08\) | \(-3.87 \pm 1.56\) | \(-0.75\) |
| 102  | 213.4708   | 52.5790     | 0.860 | 8.23\*           | 45.01 | \(-0.91 \pm 1.04\) | \(2.51 \pm 1.03\) | \(2.44\) |
| 118  | 213.5533   | 52.5358     | 0.714 | 8.48\*           | 45.12 | \(-0.90 \pm 2.92\) | \(-0.48 \pm 0.28\) | \(-0.99\) |

Notes.

\( a \) Single-epoch masses are identified by * and assumed to have an error of 0.4 dex.

\( b \) The S/N is calculated accounting for the JAVELIN lag sign; if the lag is positive, the S/N is positive, and if the lag is negative, the S/N is negative.

(This table is available in its entirety in machine-readable form.)

3.4. Criteria

We require additional tests to identify if our computed lag is statistically significant. One of the tools on which we rely is the maximum cross-correlation coefficient, \( r_{\text{max}} \), as a measure of the correlation between the g and i light curves. Visual inspection of the g and i light curves and computed lag probability distributions revealed that a threshold of \( r_{\text{max}} > 0.4 \) can eliminate noncorrelated light curves. Another tool used to identify the significance of the main peak is the fraction of the probability distribution that lies within the primary peak, hereafter referred to as “\( f_{\text{peak}} \)” We define \( f_{\text{peak}} \) as the ratio of the weighted CCCD between the local minima used in the lag calculation to the prior-weighted CCCD across the full \( \pm 100 \) day delay range. We accept only peaks that carry more than 75% of the total posterior probability (\( f_{\text{peak}} > 0.75 \)) to obtain a sample of well-measured lags from our quasar sample. We also want to avoid lags that are contaminated by BLR emission lines, as discussed above in Section 3.3.2. We thus exclude any objects with emission line contamination greater than 12.5%.

In summary, our criteria for accepting a lag as a “well-defined” lag are as follows:

1. \( r_{\text{max}} > 0.4 \): minimum cross-correlation to consider that corresponds to physical reverberation.
2. \( f_{\text{peak}} > 75\% \): threshold to reject ambiguous lags with significant support for competing aliases.
3. \( f_{\text{BLR}} < 12.5\% \): minimum broad-line contribution in both g and i photometric light curves.

Our final lag sample is reported in Table 1 for the first 10 of the 95 quasars that satisfy the above criteria. We also report redshifts (Shen et al. 2015), \( R M M_{BH} \) and single-epoch \( M_{BH} \) from Grier et al. (2017), \( \lambda L_{\lambda,3300} \) (Shen et al. 2015), and the observed-frame lag and uncertainties using both ICCF and JAVELIN.

4. Lag Reliability

The JAVELIN method produces a total of 95 well-defined lags that satisfy the reliability criteria defined in Section 3.4. From the well-defined sample of 95 continuum lags, we also construct a subsample of 33 “high-S/N” lags that are 2\( \sigma \) different from zero: S/N(\( \tau_{JAVELIN} \)) \( \geq 2 \), in addition to meeting the criteria listed in Section 3.4. Summarizing, we use the following definitions for our main sample of well-defined lags and the subsample of high-S/N lags:

1. well-defined lags: \( r_{\text{max}} > 0.4 \), \( f_{\text{peak}} > 75\% \), and \( f_{\text{BLR}} < 12.5\% \);
2. high-S/N lags: \( r_{\text{max}} > 0.4 \), \( f_{\text{peak}} > 75\% \), \( f_{\text{BLR}} < 12.5\% \), and S/N(\( \tau_{JAVELIN} \)) \( \geq 2 \).

Due to the limits in the SDSS-RM survey, our measured lags could impose selection bias. For example, the high-S/N lag sample includes only larger lags, while the well-defined lag sample may be more representative of the broader quasar population. We will discuss this point in more detail in Appendix A.

One of the difficulties in RM, particularly for monitoring surveys such as SDSS-RM with relatively sparse cadence and nonnegligible flux uncertainties, is knowing if there is genuine reverberation rather than a false detection caused by a chance similarity between light curves. Chance similarities would create an equal number of positive and negative lags, while reverberation would produce only positive lags, with some negative lags due to the noise or sampling properties of light curves. We investigate this issue with the set of plots presented in Figure 6. Our lag-finding analysis and well-defined lag criteria include no explicit or implicit preference for a positive lag from the g to i bands. The high-S/N sample has 33 positive lags and only five negative JAVELIN lags, indicating that most objects have genuine reverberation with a false-positive rate (i.e., ratio of negative to positive lag) of only 15%. The well-defined lag sample has 68 positive lags and 27 negative lags, similarly showing a significant excess of positive lags. The larger number of negative lags in the well-defined sample is expected from the broad lag CCCDs of many of the quasars.

We compare our two lag methodologies, ICCF and JAVELIN, in Figure 7. Most sources have differences between their ICCF and JAVELIN lags indicating that the ICCF uncertainties are overestimated: \( |(\tau_{\text{ICCF}} - \tau_{\text{JAVELIN}})|/\sigma_{\text{ICCF}} = 1.29 \) and \( |(\tau_{\text{ICCF}} - \tau_{\text{JAVELIN}})|/\sigma_{\text{JAVELIN}} = 0.41 \). When comparing the two methodologies, we note that JAVELIN presents an empirically

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motivated model for interpolating the light curve by explicitly assuming that the power spectral density is a DRW model, while implicitly assuming a prior that the two light curves are reverberating. The ICCF does not make this assumption and instead linearly interpolates between measurements to describe the light curve. The broad agreement between JAVELIN and ICCF lags is expected, given our relatively short observational uncertainties in the light curve by 3%, the JAVELIN results change dramatically and become consistent with the ICCF lag. Due to this object’s small error and the more than 3σ difference from the JAVELIN lag estimate, we reject this object from our sample.

Table 2 presents a comparison of our SDSS-RM study with other multi-object continuum lag surveys. Our study’s largest advantage is the availability of spectroscopic RM observations and resulting $M_{\text{BH}}$ measurements, enabling a comparison of disk size with black hole mass. Further comparison of our measured accretion disk properties with previous work is presented in Section 5.1.

5. Discussion

The photometric lags measured from SDSS-RM can be employed to measure accretion disk sizes across a wide range of quasar properties. We use the SS73 model as expressed in Equation (1) as a starting framework, comparing our measured lags to the expectations of the analytic thin-disk model.

We follow a Bayesian approach and fit accretion disk parameters using the full set of well-defined lags. Although many of these lags have large error bars and are consistent with zero, their distribution still carries valuable information. Appendix A also represents results from fitting only the high-
S/N lags, demonstrating that restricting to positive lags results in biased accretion disk fits.

We use the Bayesian framework implemented in the software package PyMC3 (Salvatier et al. 2016)13 to fit the accretion disk parameters. To sample the posterior, we provide disk parameter priors as a normal distribution centered at the expectation from the SS73 model. We sample our MCMC fit with 40,000 steps, discard the first 20,000 steps as a burn-in phase, and explicitly check the Gelman–Rubin statistics (Gelman & Rubin 1992) for a convergence diagnostic. All of the lags are reported in the observed frame (i.e., $\tau_{\text{obs}}$) as we account for the effects of wavelength redshift and time dilation in our analysis.

5.1. Disk Normalization

We start with the SS73 model presented in Equation (1) and compute each object’s individual accretion disk size $\tau_0$ following the equation for the SS73 model observed-frame lag $\tau$:

$$\tau_{\text{SS73}} = \tau_0(1 + z)^{1/3} \left[ \left( \frac{\lambda_i}{9000 \, \text{Å}} \right)^{4/3} - \left( \frac{\lambda_g}{9000 \, \text{Å}} \right)^{4/3} \right].$$

We normalize the wavelength by $\lambda_0 = \lambda/9000$ Å because it was found to minimize the correlation between the best-fit $\tau_0$ and wavelength scaling $\beta$ in Section 5.2.

For simplicity, we refer to each of the measured $C\tau$ and model-predicted $C\tau_{\text{SS73}}$ as a “disk size.” More precisely, these quantities are the relative distances corresponding to the differences between the characteristic lags from each waveband.

The analytic disk normalization $\tau_0$ is equal to

$$\tau_0 = \frac{1}{\varepsilon} \left( \frac{45 \, G}{16 \pi^6 h_p^2 \, c^2} \right)^{1/3} X^{4/3} (9000 \, \text{Å})^{4/3}$$

$$\times \left( \frac{C_{\text{bol}}}{\eta \, c^2} \right)^{1/3} M_{\text{BH}}^{1/3} \lambda_{5000}^{-1/3}. \tag{5}$$

Here $M_{\text{BH}}$ represents the black hole mass from RM (Grier et al. 2017) and single-epoch measurements (Shen et al. 2016a). When both RM and single-epoch masses are available for a quasar, we use the RM mass. We compute the $L_{\text{bol}}$ using a bolometric luminosity correction $C_{\text{bol}} = 5.15$ from Richards et al. (2006) for $\lambda_{5000}$ as $L_{\text{bol}} = C_{\text{bol}} \lambda_{5000}$. The quantity $X$ accounts for the relatively broad width of blackbody radiation causing the response at a given wavelength to arise from a range of radii in the disk, including smaller radii, where the blackbody radiation is proportional to $T$ on the Rayleigh–Jeans tail of the blackbody emission, and larger radii, where the increasing disk surface area is offset by the exponential Wien cutoff. Given a $T(r)$ profile and a wavelength $\lambda$, the observed mean delay is $\tau = \kappa(\lambda)/c$, where $\kappa = hc/XT(r(\lambda))$. We follow previous work (Fausnaugh et al. 2016; Mudd et al. 2018) and calculate $X$ by assuming that $r(\lambda)$ is the flux-weighted mean radius for emission at $\lambda$ from a face-on disk of pure blackbody emission with $r(\tau) \propto r^{-3/4}$, which yields $X = 2.49$. Comptonization and other radiative transfer effects may also affect the disk emission profile (e.g., Davis et al. 2005; Slone & Netzer 2012), potentially making $X$ a function of radius (or wavelength). We adopt the global blackbody assumption of $X = 2.49$ as a point of comparison for comparing to the SS73 model, noting that larger or smaller continuum lags may result from non-blackbody radiative transfer effects in addition to structural changes in the SS73 model.

In the following analysis, we adopt $\eta = 0.1$, $C_{\text{bol}} = 5.15$, and $X = 2.49$ when we plot the SS73 model in Figures 9, 12, 13, and 14.

Figure 8 shows a comparison of the observed lags $\tau_{\text{obs}}$ with the analytic model lags $\tau_{\text{SS73}}$ calculated from Equations (4) and (5). On average, the observed disk sizes are consistent with

### Table 2: Comparison with Other Multi-object Continuum Lag Surveys

| Survey       | Lags | Epochs | Cadence | Duration | Bands | RM $M_{\text{BH}}$ |
|--------------|------|--------|---------|----------|-------|-------------------|
| Pan-STARRS   | 39   | 373    | 3 day   | 3.3 yr   | $g, r, i, z$ | No                |
| OzDES        | 15   | 30     | 7 days  | 1 season (180 days) | $g, r, i, z$ | No               |
| SDSS-RM      | 95   | 83     | 4 days  | 1 season (180 days) | $g, i$ | Yes              |

Notes.
1. The number of reported lags for the main sample in Jiang et al. (2017a) and Mudd et al. (2018).
2. Median number of total epochs per band.
3. Time-domain spectroscopic coverage available for RM $M_{\text{BH}}$ measurements.
the SS73 model expectation (including errors on $M_{\text{BH}}$ and $M$). However, there is large scatter, with only 36% of the observed well-defined lags lying within 1σ of the model lags. The large scatter might indicate that the JAVELIN lag uncertainties are underestimated or that there are additional important parameters missing from Equations (4) and (5), such as nonuniform efficiency or orientation. We discuss this issue further in Section 5.3.

We perform an initial fit to the disk size by first allowing the normalization $\tau_0$ to be the only free parameter and fixing $\beta = 4/3$. Then, MCMC samples the posterior distribution of $\tau_0$. Fitting only the disk normalization based on all of the observed quasar lags in the well-defined sample results in a best-fit disk normalization $\tau_0|_{\beta=4/3} = 5.21^{+0.29}_{-0.29}$ days. This is consistent within 1σ with the SS73 disk normalization, $\langle \tau_0 \rangle = 4.78$ days, computed using Equation (5) for the mean $\langle M_{\text{BH}} \rangle = 8.19 M_\odot$ and $\langle \lambda_{3000} \rangle = 44.47$ of our sample. We compare our results to those from microlensing (Morgan et al. 2010) and find that our lags are three to four times larger than theirs, but this can be attributed to the fact that they use $X = 1$ in Equation (5), so inflating the SS73 disks of Morgan et al. (2010) by $X = 2.49$ will give consistent results with the SS73 expectation (see also Tie & Kochanek et al. 2018). In contrast, Jiang et al. (2017a) found lags that are about two to three times larger than SS73. However, the Jiang et al. (2017a) lag sample, by including only significant lags, is biased toward larger lags and thus larger disk sizes. The implication of the bias is less apparent in the recent work by Mudd et al. (2018), where they reported consistent lags with SS73. Our measured accretion disk sizes are similar to those found by Mudd et al. (2018), in that both of our results are broadly consistent with the SS73 model. However, our study has the additional advantage of $M_{\text{BH}}$ estimates from spectroscopic RM, which we use in Section 5.3 to model accretion disk size as a function of black hole mass and luminosity. In Appendix A, we discuss the effects of observational bias on lag measurements of multi-object quasar samples. Mixed results are reported for more local quasars; e.g., some report lags that are too big (Edelson et al. 2015, 2017; Fausnaugh et al. 2016, 2018), and some report lags that are close to the SS73 expectation (McHardy et al. 2018). These results may be due to local objects from the NGC sample probing the biased tail of the quasar distribution.

5.2. Color Profile

The SS73 accretion disk model predicts a disk structure of $T(R) \propto R^{1.4}$. We measure this temperature profile using the wavelength in Equation (6) with a disk size that is characterized by a disk normalization $\tau_0$, wavelength scaling $\beta$, and quasar redshift $z$. In this context, the observed continuum lags are described by

$$\tau_{\text{obs}} = \tau_0(1 + z)^{(1-\beta)} \left[ \frac{\lambda_1}{9000 \ \text{Å}} \right]^{\beta} - \left[ \frac{\lambda_2}{9000 \ \text{Å}} \right]^{\beta}. \quad (6)$$

Although we are only limited to the $g$ and $i$ bands in this work, the redshift range of our quasars ($0.116 < z < 1.128$) provides a broad range of rest-frame wavelengths to test $\beta$, with the best-fit disk size and color profile shown in Figure 9. The best-fit parameters and errors are determined from the posterior distributions of the MCMC nonlinear regression. We assume the likelihood as a normal distribution, $\mathcal{N}$, centered at the observed lags and lag errors as a standard deviation

$$P(\theta|X) = \mathcal{N}(\tau_{\text{model}}, \tau_{\text{obs}}, \sigma_{\text{lag}}). \quad (7)$$

Posterior distributions are shown in Figure 10; we find $\tau_0 = 5.38^{+0.43}_{-0.34}$ days and $\beta = 1.28^{+0.41}_{-0.39}$.

Comparing the best-fit $\tau_0$ and color profile $\beta$ to the SS73 model indicates that the best-fit values are consistent with the SS73 expectation for our sample of mean $M_{\text{BH}}$ and $L_{\text{bol}}\ldots$
Our best-fit color profile $\beta$ is also consistent within 1$\sigma$ with previous results by Fausnaugh et al. (2016) and Mudd et al. (2018); further comparison with Fausnaugh et al. (2016) requires multiband observations, as we are only comparing the $g$ and $i$ bands here. For the remaining portion of this work, we will fix $\beta$ to $4/3$ in order to focus on the accretion disk connections to $M_{\text{BH}}$ and accretion rate.

5.3. Connection to $M_{\text{BH}}$ and $\lambda L_{\lambda 3000}$

Here we examine if our measured continuum lags depend on $M_{\text{BH}}^{4/3}$ and $M_{\text{BH}}^{1/3}$ as indicated by the SS73 model. Our 95 quasars in the well-defined lag sample have reliable $M_{\text{BH}}$ estimates using the RM technique for 30 of the quasars and single-epoch mass measurements for the remaining 65 quasars; see Table 1. To test for connections to $M_*$, we use the observable monochromatic luminosity $\lambda L_{\lambda 3000}$ as a proxy for $M_*$, related as $\dot{M} = L_{\text{bol}}/\eta c^2$, with $L_{\text{bol}} = 5.15 \lambda L_{\lambda 3000}$. In this context, the observed continuum lags are described by

$$\tau_{\text{obs}} = \tau_0' \left( \frac{M_{\text{BH}}}{10^8 M_\odot} \right) \left( \frac{\lambda L_{\lambda 3000}}{10^{44} \text{ erg s}^{-1}} \right)^{\gamma} (1 + z)^{1-\beta} \times \left[ \frac{\lambda_v}{9000 \text{ Å}} \right]^{\delta} \left[ \frac{\lambda_{\text{bol}}}{9000 \text{ Å}} \right]^{\delta}.$$  

We perform a new nonlinear MCMC regression fit for $\tau_0'$, $\gamma$, and $\delta$. Here $\tau_0'$ has a slightly different form from the previous disk normalization due to different powers in mass and luminosity (i.e., $\tau_0' = \tau_0/M_{\text{BH}}^{4/3} \lambda L_{\lambda 3000}$). We fix $\beta = 4/3$ in Equation (8) and incorporate the measurement uncertainties in $M_{\text{BH}}$ reported by Grier et al. (2017). The uncertainties in RM $M_{\text{BH}}$ include a 0.16 dex intrinsic scatter, while for single-epoch $M_{\text{BH}}$ estimates,
we assume a 0.4 dex intrinsic scatter (Vestergaard & Peterson 2006; Shen et al. 2016b). We also incorporate the measurement uncertainties while fitting to the observed $\lambda L_{3000}$. However, the SS73 model predicts disk size as a function of $M$ rather than $\lambda L_{3000}$ and there is a large scatter between the observed luminosity and accretion rate due to uncertainties in bolometric correction and radiative efficiency. This might effectively lead to a larger scatter in the fit, which we measure in the regression fit using an excess dispersion parameter $\sigma$.

The result of our three-parameter disk model to the well-defined sample is illustrated in Figures 11 and 12.

With the disk size parameterized as $\tau_0$, $M_{BH}$, and $\lambda L_{3000}$ (Equation (8)), we find the best-fit $\tau_0 = 4.16^{+1.37}_{-1.37}$ days, $\gamma = 0.31^{+0.21}_{-0.20}$, and $\delta = 0.09^{+0.16}_{-0.18}$. Both the $\gamma$ and $\delta$ parameters are poorly constrained, although the mass dependence is $>1\sigma$ different from zero and fully consistent with the SS73 expectation $\gamma = 1/3$. Our fit indicates that luminosity $\lambda L_{3000}$ on the other hand, is less necessary for the fit, differing from the SS73 expectation by $1.5\sigma$. A more accurate measurement of $M_{BH}$ could improve the consistency (i.e., in Equation (1)).

Our best-fit parameters include an intrinsic excess dispersion of 2.8 days. This could indicate that the JAVELIN lag errors are underestimated, although the good agreement with ICCF lags in Figure 7 suggests that this is unlikely. Alternatively, individual quasars may have diverse disk profile profiles, with a range of orientation and/or radiative transfer effects that change the X factor in our parameterization (Hall et al. 2018). Some quasars may also have significant continuum emission from a diffuse BLR component, making the measured interband lags differ from pure accretion disk continuum emission (Cackett et al. 2018; Edelson et al. 2019). A nonuniform bolometric correction or radiative efficiency might also lead to scatter in our best-fit disk size as a function of monochromatic luminosity (Equation (6)), although this would have to be as large as 1.8 dex to explain the entirety of the excess scatter measured of $d\tau/\tau = 1.35$ in our regression fit. Finally, it is possible that the SS73 model is a good average description for quasar disks, even as individual objects have large variations in their disk structure not captured by the model.

6. Summary

We have used continuum RM to study the accretion disks of 222 quasars from the SDSS-RM survey. The selected sample has the advantage of reliable black hole mass measurements from the first year of the SDSS-RM monitoring program (Grier et al. 2017). In this work, we used photometric continuum light curves in the $g$ and $i$ bands to study the accretion disk size and structure of quasars.

We used JAVELIN to compute lags between $g$- and $i$-band light curves for our 222 quasars. We applied several different significance criteria to obtain a subset of 95 well-defined continuum lags.

Purely comparing our observed lags to those expected from the SS73 model, we find a mean deviation of 0.9 days larger than the SS73 expectation, with 36% of the well-defined lags consistent within $\pm 1\sigma$ of the SS73 model expectation. We perform nonlinear MCMC regression to fit our observed lags and compare them to the standard SS73 model. Our findings are as follows.

1. Disk size. Our best-fit disk normalization is consistent with the theoretical value from SS73 within $1.5\sigma$. This is in contrast to previous works, possibly due to observational bias (as discussed in Appendix A).

2. Color profile. We find a wavelength scaling $\beta = 1.28^{+0.1}_{-0.3}$, consistent with the SS73 expectation (i.e., $\beta = 4/3$).

3. Mass and luminosity dependence. We assume a disk size $\tau \propto \lambda L_{3000}$ and find a best-fit mass dependence $\gamma = 0.31^{+0.21}_{-0.20}$, consistent with expectations from SS73 (i.e., $1/3$).
The best-fit $\lambda L_{\lambda,3000}$ dependence is $\delta = 0.09^{+0.16}_{-0.16}$, 1.4$\sigma$ consistent with the SS73 expectation but also $<1\sigma$ consistent with no correlation between disk size and luminosity. Our fits have a large excess dispersion of 2.8 days, indicating a diversity of radiative efficiency, disk emission profiles, and/or disk structure in individual quasars.

Our new measurements represent a large advance over previous work. The 95 SDSS-RM quasars with our new continuum lags and previous broad-line lags (Grier et al. 2017) represent a factor of $\sim 5$ increase over previous samples and also expand the sample of accretion disk size and black hole mass measurements by an order of magnitude in redshift, mass, and luminosity. Our measured disk sizes are, on average, consistent with the SS73 analytic thin-disk model. But we also find a large range of smaller and larger disk sizes in excess of the measurement uncertainties. This motivates future work to better measure bolometric luminosity and radiative efficiency (i.e., black hole spin) alongside accretion disk sizes.

Our work also advances the methodology for accretion disk size measurements from similar “industrial-scale” multi-object reverberation projects beyond SDSS-RM. In particular, we advocate a Bayesian approach to the full sample of well-defined lag measurements, rather than restricting the analysis to a set of high-S/N lags that are biased by limitations in survey cadence. The SDSS-RM is planned to continue in the 2020s with a factor of 5 increase in survey area as part of the SDSS-V black hole mapper project (Ivezić et al. 2017; Kollmeier et al. 2017). The Large Synoptic Survey Telescope will usher in an entirely new era of time-domain quasar studies, making continuum RM possible for thousands of quasars in its deep drilling fields.

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Appendix A

Selection Bias

We take a Bayesian approach in Section 5 and fit all 95 quasars with well-defined lags (see Section 3.4), including those that are consistent with zero lag. However, if we instead fit only the high-S/N lags (well defined and lag S/N > 2$\sigma$), we find disks that are $\sim 3.2$ times larger than SS73 and a nearly flat color profile, $\beta = 0.4$, shown in Figure 13. Additionally, we test for well-defined and positive lags and find disks that are $\sim 2.5$ times larger than expected by SS73; see Figure 14.

The high-S/N sample is biased to large lags, as the SDSS-RM cadence (averaging 4 days) sets a minimum detectable lag. This biases the disk fits to large values. A similar bias is likely to affect the main sample in Jiang et al. (2017a), as they used only positive lags in their fits. We reproduce the same qualitative effects if we limit our sample to only positive lags; see Figure 14.

Our larger well-defined lag sample, on the other hand, is not biased to large lags. Although the sample includes many lags that are formally consistent with zero, the lags are more likely to be positive than negative, as shown in Figure 6. This indicates that the lags are likely the result of genuine reverberation but are just smaller than detectable by the SDSS-RM cadence (average of 4 days). In other words, the well-defined sample includes many lags that have poor S/N but are constrained to be small. It is important to include such lags in the accretion disk fits to avoid a bias to large disk sizes.
Appendix B
Increased Short-timescale Variability

The SDSS-RM time-monitoring observations (see Section 2) are limited by somewhat sparsely sampled data with a median cadence of 4 days. Although the expected quasar variability on such short timescales is relatively low (Mushotzky et al. 2011; MacLeod et al. 2012), our measured lags are fundamentally limited by the observational cadence. Here we validate that quasar fluctuations on timescales shorter than our observation cadence do not affect our lag measurements by constructing synthetic light curves that have extreme variability in between each measured point.

To construct our new synthetic light curves, we take each consecutive measured flux pair \( f(t_i), f(t_{i+1}) \) and randomly select an interpoint expectation flux \( f(t_{i+1/2}) \), where \( t_{i+1/2} = t_i + (t_{i+1} - t_i)/2 \), from the JAVELIN DRW model normal distribution. We then increase (or decrease) each randomly selected flux to a new flux \( f(t_{i+1/2}) + \delta f \), where \( \delta f \equiv 1/2 \).
Figure 15. Top: synthetic continuum g- (blue) and i-band (red) light curves with increased short-timescale variability in between observations. The interpoint variability was increased between each consecutive pair of observations. The interpoint flux was randomly selected between consecutive observations using the JAVELIN DRW model normal distribution increased (decreased) by half the difference of the observation pair. Bottom: we find the same lag with both JAVELIN and ICCF methods.

\[(f(t_{i+1}) - f(t_i))\], i.e., varying by half the difference between consecutive pairs of measured fluxes. This is equivalent to a short-timescale variability PSD of \(\alpha = -1\): an extreme variability case compared to a DRW (\(\alpha = -2\)) and the low measured short-timescale variability of \(\alpha \approx -3\) (Mushotzky et al. 2011). We also perturbed the new interpoint flux by the average flux uncertainty of the measured surrounding flux pair. The final synthetic light curve is then the combination of both the measured light curve and the new interpoint fluxes.

We build synthetic light curves for all of our targets in the well-defined sample and use JAVELIN to measure lags as described in Section 3.3. We find that the synthetic light curves have lags that are statistically consistent with the original lags measured from the observed light curves. Figure 15 illustrates the synthetic light curve with extreme short-timescale variability for RM 267 (the same target as Figure 4) and the measured JAVELIN and ICCF lag probability distributions.

**References**

Abramowicz, M. A., Czerny, B., Lasota, J. P., & Szuszkiewicz, E. 1988, *ApJ*, 332, 646
Alard, C. 2000, *A&AS*, 144, 363
Almaini, O., Lawrence, A., Shanks, T., et al. 2000, *MNRAS*, 315, 325
Aune, S., Boulade, O., Charlot, X., et al. 2003, *Proc. SPIE*, 4841, 513
Bentz, M. C., Denney, K. D., Grier, C. J., et al. 2013, *ApJ*, 767, 149
Blackburne, J. A., Pooley, D., Rappaport, S., & Schechter, P. L. 2011, *ApJ*, 729, 34
Blandford, R. D., & McKee, C. F. 1982, *ApJ*, 255, 419
Bolton, A. S., Schlegel, D. J., Aubourg, E., et al. 2012, *AJ*, 144, 144
Cackett, E. M., Chaing, C.-Y., McHardy, I., et al. 2018, *ApJ*, 857, 53
Cackett, E. M., Horne, K., & Winkler, H. 2007, *MNRAS*, 380, 669
Collier, S. J., Horne, K., Kaspi, S., et al. 1998, *ApJ*, 500, 162
Davis, S. W., Blues, O. M., Hubeny, I., & Turner, N. J. 2005, *ApJ*, 621, 372
Dawson, K. S., Schlegel, D. J., Ahn, C. P., et al. 2013, *AJ*, 145, 10
Denney, K. D., Horne, K., Brandt, W. N., et al. 2016a, *ApJ*, 833, 33
Denney, K. D., Horne, K., Shen, Y., et al. 2016b, *ApJ*, 224, 14
Destroches, L.-B., Greene, J. E., Ho, L. C., et al. 2009, *ApJ*, 698, 1515
Dexter, J., & Agol, E. 2011, *ApJ*, 727, 24
Doi, M., Tanaka, M., Fukugita, M., et al. 2010, *AJ*, 139, 1628
Du, P., Hu, C., Lu, K.-X., et al. 2015, *ApJ*, 806, 22
Edelson, R., Gelbord, J., Cackett, E., et al. 2017, *ApJ*, 840, 41
Edelson, R., Gelbord, J., Cackett, E., et al. 2019, *ApJ*, 870, 123
Edelson, R., Gelbord, J. M., Horne, K., et al. 2015, *ApJ*, 806, 129
Eisenstein, D. J., Weinberg, D. H., Agol, E., et al. 2011, *ApJ*, 142, 72
Elitzur, M., Ho, L. C., Trump, J. R., et al. 2014, *MNRAS*, 438, 3340
Fausnaugh, M. M., Denney, K. D., Barth, A. J., et al. 2016, *ApJ*, 821, 56
Fausnaugh, M. M., Grier, C. J., Bentz, M. C., et al. 2017, *ApJ*, 840, 97
Fausnaugh, M. M., Starkey, D. A., Horne, K., et al. 2018, *ApJ*, 854, 107
Fukugita, M., Ichikawa, T., Gunn, J. E., et al. 1996, *AJ*, 111, 1748
Gaskell, C. M., & Peterson, B. M. 1987, *ApJS*, 65, 1
Gaskell, C. M., & Spark, L. S. 1986, *ApJ*, 305, 175
Gehrels, N., Chincarini, G., Giommi, P., et al. 2004, *ApJ*, 611, 1005
Gelman, A., & Rubin, D. B. 1992, *StatSci*, 7, 457
Grier, C. J., Brandt, W. N., Hall, P. B., et al. 2016, *ApJ*, 824, 130
Grier, C. J., Trump, J. R., Shen, Y., et al. 2017, *ApJ*, 851, 21

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Gunn, J. E., Siegmund, W. A., Mannery, E. J., et al. 2006, AJ, 131, 2332G
Ho, L. C. 2008, ARA&A, 46, 475
Ivezić, Z., Connolly, A. J., & Jurić, M. 2017, arXiv:1612.04772
Jiang, Y.-F., Green, P. J., Greene, J. E., et al. 2017a, ApJ, 836, 186
Jiang, Y.-F., Stone, J., & Davis, S. W. 2014, ApJ, 796, 106
Jiang, Y.-F., Stone, J., & Davis, S. W. 2017b, arXiv:1709.02845
Kaspi, S., Brandt, W. N., Maoz, D., et al. 2007, ApJ, 659, 997
Kelly, B. C., Bechtold, J., & Siemiginowska, A. 2009, ApJ, 698, 895
Kollmeier, J. A., Zasowski, G., Rix, H.-W., et al. 2017, arXiv:1711.03234
Kozłowski, S., Kochanek, C. S., Udalski, A., et al. 2010, ApJ, 708, 927
Krolik, J. H., Horne, K., Kallman, T. R., et al. 1991, ApJ, 371, 541K
Laor, A., & Davis, S. W. 2014, MNRAS, 438, 3024
Lawrence, A. 2005, MNRAS, 363, 57L
Li, J., Shen, Y., Brandt, W. N., et al. 2019, ApJ, submitted
Li, J., Shen, Y., Horne, K., et al. 2017, ApJ, 846, 79
Luo, B., Brandt, W. N., Hall, P. B., et al. 2015, ApJ, 805, 122
MacLeod, C. L., Ivezić, Z., Kochanek, C. S., et al. 2010, ApJ, 721, 1014
MacLeod, C. L., Ivezić, Z., Sesar, B., et al. 2012, ApJ, 753, 106
Matsuoka, Y., Strauss, M. A., Shen, Y., et al. 2015, ApJ, 811, 91
McHardy, I. M., Cameron, D. T., Dwelly, T., et al. 2014, MNRAS, 444, 1496
McHardy, I. M., Connolly, S. D., Horne, K., et al. 2018, MNRAS, 480, 2601
McKinney, J. C., Tchekhovskoy, A., Sadowski, A., & Narayan, R. 2014, MNRAS, 441, 3177
Morgan, C. W., Kochanek, C. S., Morgan, N. D., & Falco, E. E. 2010, ApJ, 712, 1129
Motta, V., Mediavilla, E., & Rojas, K. 2017, ApJ, 835, 132
Mudd, D., Martini, P., Zu, Y., et al. 2018, ApJ, 862, 123
Mushtaq, R. F., Edelson, R., Baumgartner, W., & Gaudi, P. 2011, ApJ, 743, 12
Narayan, R., & McClintock, J. E. 2008, NewAR, 51, 733
Narayan, R., & Yi, I. 1994, ApJ, 428, 13
Peterson, B. M. 2004, IAUS, 222, 15P
Richards, G. T., Lacy, M., Storrie-Lombardi, L. J., et al. 2006, ApJS, 166, 470
Sadowski, A., Lasota, J.-P., Abramowicz, M. A., & Narayan, R. 2016, MNRAS, 456, 3915S
Sadowski, A., Narayan, R., McKinney, J. C., & Tchekhovskoy, A. 2014, MNRAS, 439, 503
Salvatier, J., Wiecki, T. V., & Fonnesbeck, C. 2016, Peerj Computer Science, 2, e65
Sergeev, S. G., Doroshenko, V. T., Golubinskiy, Y. V., Merkulova, N. I., & Sergeeva, E. A. 2005, ApJ, 622, 129
Shakura, N. I., & Sunyaev, R. A. 1973, A&A, 24, 337
Shappee, B. J., Prieto, J. L., Grange, D., et al. 2014, ApJ, 788, 48
Shen, Y., Brandt, W. N., Dawson, K. S., et al. 2015, ApJ, 816, 4
Shen, Y., Brandt, W. N., Richards, G. T., et al. 2016a, ApJ, 831, 7
Shen, Y., Hall, P., Keith, H., et al. 2019, ApJS, 241, 34
Shen, Y., Horne, K., Grier, C. J., et al. 2016b, ApJ, 818, 30
Slone, O., & Netzer, H. 2012, MNRAS, 426, 656
Smee, S. A., Gunn, J. E., Uomoto, A., et al. 2013, AJ, 146, 32
Starkey, D. A., Horne, K., Fausnaugh, M. M., et al. 2017, ApJ, 835, 65
Starkey, D. A., Horne, K., & Villforth, C. 2016, MNRAS, 456, 1960
Sun, M., Grier, C. J., & Peterson, B. M. 2018a, PyCCF: Python Cross Correlation Function for reverberation mapping studies, Astrophysics Source Code Library, ascl:1805.032
Sun, M., Trump, J. R., Shen, Y., et al. 2015, ApJ, 811, 42
Sun, M., Xue, Y., Richards, G. T., et al. 2018b, ApJ, 854, 128
Sun, M., Xue, Y., Trump, J. R., et al. 2019, MNRAS, 482, 2788
Sun, M., Xue, Y., Wang, J., et al. 2018c, ApJ, 866, 74
Tie, S. S., Kochanek, C. S., et al. 2018, MNRAS, 473, 80
Trump, J. R., Impey, C. D., Kelly, B. C., et al. 2011, ApJ, 733, 60
Vestergaard, M., & Peterson, B. M. 2006, ApJ, 641, 689
Wanders, I., Peterson, B. M., Allin, D., et al. 1997, ApJS, 113, 69
White, R. J., & Peterson, B. M. 1994, PASP, 106, 879
Williams, G. G., Olszewski, E., Lesser, M. P., & Burge, J. H. 2004, Proc. SPIE, 5492, 787
Zu, Y., Kochanek, C. S., Peterson, B. M., et al. 2011, ApJ, 735, 80