Optimal parameter estimation of Pauli channels

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Received 12 January 2012, in final form 15 May 2012
Published 15 June 2012
Online at stacks.iop.org/JPhysA/45/265305

Abstract

The optimal measurement configuration, i.e. the optimal input quantum state and measurement in the form of a positive operator-valued measure with two elements, is investigated in this paper for qubit and generalized Pauli channels. The channel directions are defined as the contracting directions of the channel with an arbitrary fixed basis of Pauli matrices. In the qubit Pauli channel case with known channel directions it is shown that the optimal configuration that maximizes the Fisher information of the estimated parameters consists of three von Neumann measurements together with pure input states directed to the appropriate channel directions. Extensions of these results for generalized Pauli channels are also given. Furthermore, a computationally efficient estimation method is proposed for estimating the channel directions in the qubit Pauli channel case. The efficiency of this method is illustrated by simulations: it is shown that its variance is of order 1/N, where N is the number of used measurements.

PACS numbers: 03.67.-a, 03.65.Wj

(Some figures may appear in colour only in the online journal)

1. Introduction

Accurate description of different quantum phenomena is a key issue in the potential use of these phenomena in modern IT technologies: communication, security, quantum computers, etc. The parameter estimation of quantum channels, which is commonly called quantum process tomography [6], plays a major role in quantum information processing. The direct way of quantum process tomography is performed by sending known quantum systems into the channel, and then estimating the output state. Experimental design is necessary if one wants to get efficient process tomography that consists of selecting the optimal input state, optimal measurement of the output state and an efficient estimator of channel from the measurement.
data. In quantum mechanics, the measurement has a probabilistic nature [12, 14]; therefore, many identical copies of the input quantum system are needed, and an estimator can be constructed by using some statistical considerations.

The field of quantum process tomography is well established; an exhaustive description of possible tomography methods can be found in [10]. Within this field a relatively wide family of quantum channels described as Pauli channels [12] are of large theoretical and practical importance. The tomography of Pauli channels has been covered by a huge literature but—because of the difficulty of the topic—the papers mostly deal with special cases. For example, in [17] the depolarization channel is examined. In the case of this one-parameter family of channels, the optimal measurement setup can be easily found. The optimal measurement is made via coupling an ancilla system, which is a common method in channel tomography [2, 9, 10, 16], because the tomographic efficiency can be improved via entanglement. At the same time, this is not necessary true, in general [4], so direct methods are widely spread too.

If one wants to find an optimal estimation method, the common solution is that some kind of Fisher information is maximized, which, by the Cramer–Rao inequality, results in a minimal variance. A similar problem appears in [7], where the discrimination of two different Pauli channels is investigated. While the efficiency of the discrimination requires another quantity, the optimal estimation method is the same as in standard tomography problems. Many recent experimental results have also been proposed that provide the optimal estimation of the Pauli channels in practice [3, 5, 19].

Each estimation method is constrained by the fact that quantum channels must be completely positive and trace-preserving (CPTP); otherwise, their description is physically not meaningful. In case one has obtained a non-physical map, it can be used to construct one that is physical [20]. In the case of a large number of measurements, however, the law of large numbers ensures that the estimated values will be close to the real channel parameters, which are obviously physically possible.

While most papers deal with the qubit and one-parameter case, there are some exceptions. There are some papers investigating the estimation of multi-parameter channels [2, 19, 13], and the multidimensional case also appears in [8] and [11]. Although there are alternative ways for the generalization of Pauli channels to the multidimensional case, the above cited papers use a given basis and represent the Pauli channels by a contraction in every direction. The paper [15] proposes another way for the generalization of Pauli channels: the state space is divided on complementary subalgebras, and the contraction is applied on these subalgebras. This way the complete positiveness of mapping can be easily derived. This latter approach is followed here when dealing with the generalized Pauli channels.

The aim of this work is to find the optimal measurement configuration consisting of an optimal quantum input state and a positive operator-valued measure (POVM) that provide a statistically optimal estimate of the channel parameters. For the sake of simplicity, in this work we will only use POVMs with two elements. An earlier engineering approach for this problem, that is called an experiment design problem in system identification, has been reported in [1], where the problem of optimal estimation of Pauli-channel parameters was investigated using convex optimization methods. In contrast to this, purely statistical methods are applied in this paper that enable us to obtain results analytically.

The paper is organized as follows. In section 2 we introduce the necessary notions to understand the rest of the paper. In sections 3 and 4 we assume that the channel directions are known and give the optimal measurement configuration for both the qubit and the generalized Pauli-channel cases. In section 5 the method for estimating the channel directions is described and analyzed. Finally, conclusions are drawn.
2. Preliminaries

2.1. Quantum states and directions

The quantum state of a finite $n$-dimensional system can be described using $\rho \in M_n(\mathbb{C})$ fulfilling the conditions

$$\rho \succeq 0, \quad \text{Tr}(\rho) = 1. \quad (1)$$

In the special case of qubits (i.e. $n = 2$), we use Bloch parametrization:

$$\rho_\theta = \frac{1}{2} \left[ \begin{array}{cc} 1 + \theta_3 & \theta_1 - i \theta_2 \\ \theta_1 + i \theta_2 & 1 - \theta_3 \end{array} \right] = \frac{1}{2} \left( I + \sum_{i=1}^3 \theta_i \sigma_i \right), \quad (2)$$

where the so-called Pauli matrices

$$I = \sigma_0 = \left[ \begin{array}{cc} 1 & 0 \\ 0 & 1 \end{array} \right], \quad \sigma_1 = \left[ \begin{array}{cc} 0 & 1 \\ 1 & 0 \end{array} \right], \quad \sigma_2 = \left[ \begin{array}{cc} 0 & -i \\ i & 0 \end{array} \right], \quad \sigma_3 = \left[ \begin{array}{cc} 1 & 0 \\ 0 & -1 \end{array} \right] \quad (3)$$

are used as a basis in $M_2(\mathbb{C})$. From (1), it follows that all states can be described with the Bloch vector $\theta = (\theta_1, \theta_2, \theta_3) \in \mathbb{R}^3$ fulfilling the condition

$$\theta_1^2 + \theta_2^2 + \theta_3^2 \leq 1. \quad (4)$$

In the $n$-dimensional case let us use an arbitrary self-adjoint orthonormal basis $v_i$ instead of Pauli matrices that results in the parametrization

$$\rho_\theta = \frac{2^{n-1}}{n} \sum_{i=0}^{2^{n-1}-1} \theta_i v_i = \frac{I}{n} + \sum_{i=1}^{2^{n-1}-1} \theta_i v_i, \quad (5)$$

where we used $v_0 = \frac{1}{\sqrt{n}} I$. Then from (1) it follows that $\theta_0 = \frac{1}{\sqrt{n}}$, but the positivity condition cannot be written in a simple, explicit form [14].

If $n = 2^k$, then we can obtain the basis from the tensor product of the Pauli matrices:

$$\frac{1}{\sqrt{n}} \otimes_{i=1}^k \sigma_{j_i}, \quad \text{where } \forall i: j_i \in \{0, 1, 2, 3\}. \quad (6)$$

Since density matrices can be described by Bloch vectors, we can consider the quantum states as elements of $\mathbb{R}^{2^n-1}$. The direction of a Bloch vector is defined as the direction of this vector in $\mathbb{R}^{2^n-1}$ with the basis of the Pauli matrices.

2.2. Quantum measurements

Quantum measurements are represented mathematically by POVMs. These can be described as a set of $n \times n$ matrices satisfying

$$\forall a : M_a \succeq 0 \quad \text{and} \quad \sum_a M_a = I. \quad (6)$$

We will use only POVMs with two elements: $\{M, I - M\}$. We can parametrize the POVM element $M$ in the same basis as we did in (5):

$$M = \sum_{i=0}^{n-1} m_i v_i = m_0 \sqrt{n} \left( \frac{I}{n} + \sum_{i=1}^{n-1} \frac{m_i}{m_0 \sqrt{n}} v_i \right) = m_0 \sqrt{n} \left( \frac{I}{n} + \sum_{i=1}^{n-1} m_i^* v_i \right). \quad (7)$$

Thus, $M$ can be represented with a Bloch vector $m = (m_0, m_1, \ldots, m_{n-1}) \in \mathbb{R}^n$. The state space of such POVMs has a similar structure as the space of quantum states (cf $\theta$ and $m^*$); the difference is that $m_0$ is not fixed. The maximal value for $m_0$ is $\sqrt{n}$; in that case we have $\forall i > 0 : m_i = 0$ and so $M = I$. An arbitrary POVM fulfills the conditions in (6) if all eigenvalues of $M$ are in the interval $[0, 1]$. If all eigenvalues are 0 or 1, then $M$ is a projection and the POVM is called a von Neumann measurement.
The probability of measuring $M$ is

$$p = \text{Tr}(\rho_0 M) = \sum_{i=0}^{n-1} m_i \theta_i = \frac{m_0}{\sqrt{n}} + \sum_{i=1}^{n-1} m_i \theta_i = \frac{m_0}{\sqrt{n}} + d,$$

where we use the orthonormality of basis ($\langle v_i, v_j \rangle = \delta_{i,j}$) and the abbreviation

$$d = \sum_{i=1}^{n-1} m_i \theta_i.$$

2.3. Quantum Pauli channels

Quantum Pauli channels form a well-known and wide family of quantum channels in the two-dimensional (qubit) case. Their usual definition implements contractions in the directions of the Pauli matrices (basis vectors), but we define it with arbitrary directions.

Let $v_i$ be a self-adjoint orthogonal basis of $M_2(\mathbb{C})$; then the effect of the channel on an input state $\rho_{\theta}$ can be described as

$$\rho_{\theta} = \frac{1}{2} \left( I + \sum_{i=1}^{3} \theta_i v_i \right) \rightarrow \mathcal{E}(\rho_{\theta}) = \frac{1}{2} \left( I + \sum_{i=1}^{3} \lambda_i \theta_i v_i \right).$$

Hence the image of the Bloch ball will be an ellipsoid with semi-axes corresponding to $v_i$ directions and with length $\lambda_i$.

A possible way of defining the generalized Pauli channels is proposed in [15]. Assume that $M_n$ contains complementary subalgebras $A_1, A_2, \ldots, A_{N_s}$, where all $A_i$ have the same dimension $s$. They are complementary in the sense that their traceless parts are orthogonal:

$$A_i - \text{Tr} A_i I \perp A_j - \text{Tr} A_j I, \quad \forall A_i \in A_i, A_j \in A_j, \quad i \neq j.$$

The above subalgebras represent now the channel directions. Then by simple dimension counting we obtain

$$(s-1)N_s = n^2 - 1.$$

In order to describe the effect of the generalized Pauli channel, the trace preserving projection $E_i : M_n \rightarrow A_i$ is defined first which is usually called conditional expectation. Then, for any input $A \in B(H)$ in the form

$$A = -\frac{(N_s - 1)\text{Tr}A}{n} + \sum_{i=1}^{N_s} E_i(A),$$

the output of the generalized Pauli channel is

$$\mathcal{E}(A) = \left( I - \sum_{i=1}^{N_s} \lambda_i \right) \frac{\text{Tr} A}{n} I + \sum_{i=1}^{N_s} \lambda_i E_i(A).$$

The conditions for CPTP property are

$$1 + n\lambda_i \geq \sum_{j} \lambda_j \geq -\frac{1}{n-1}, \quad |\lambda_i| \leq 1.$$

If the subalgebras are generated by some elements of the basis $\{v_i\}$ in (5), then the channel can be described as

$$\mathcal{E}(\rho_{\theta}) = \mathcal{E} \left( \frac{1}{n} I + \sum_{i=1}^{n-1} \theta_i v_i \right) = \frac{1}{n} I + \sum_{i=1}^{n-1} \lambda_{\pi_i} \theta_i v_i,$$

where $\pi_i = k$, if $v_i \in A_k$. 

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Example 1. The simplest decomposition of $M_4$ to complementary subalgebras is

$$A_i = \text{span}[v_0, v_i], \quad i \in \{1, 2, \ldots, 15\}.$$ 

Here we have subalgebras isomorphic to $\mathbb{C}^2$, and from (13) the effect of the channel will be $\mathcal{E}(\rho_0) = \frac{1}{4} I + \sum_{i=1}^{15} \lambda_i v_i$. 

Example 2. Another decomposition of $M_4$ to complementary subalgebras:

$$A_1 = \text{span}[I, \sigma_0 \otimes \sigma_1, \sigma_0 \otimes \sigma_2, \sigma_0 \otimes \sigma_3], \quad A_2 = \text{span}[I, \sigma_1 \otimes \sigma_0, \sigma_2 \otimes \sigma_0, \sigma_3 \otimes \sigma_0]$$

$$A_3 = \text{span}[I, \sigma_1 \otimes \sigma_1, \sigma_2 \otimes \sigma_2, \sigma_3 \otimes \sigma_2], \quad A_4 = \text{span}[I, \sigma_1 \otimes \sigma_2, \sigma_2 \otimes \sigma_3, \sigma_3 \otimes \sigma_1]$$

$$A_5 = \text{span}[I, \sigma_1 \otimes \sigma_3, \sigma_2 \otimes \sigma_3, \sigma_2 \otimes \sigma_1].$$

Let us remark that $A_1 = CI \otimes M_2$ and $A_3 = M_2 \otimes CI$ are isomorphic to $M_2$, while $A_1, A_4$ and $A_5$ are isomorphic to $\mathbb{C}^4$. Formula (13) means heuristically that we make contraction with constant $\lambda_j$ ($j = 1, 2, \ldots, 5$) in the $A_j$ direction, e.g., we multiply with $\lambda_j$ if $v_i = \sigma_0 \otimes \sigma_1$, or $v_i = \sigma_0 \otimes \sigma_2$, or $v_i = \sigma_0 \otimes \sigma_3$.

2.4. Parameter estimation

If we want to estimate the channel parameters, first we send an input state ($\rho_0$) through the channel, then perform a measurement ($M, I - M$) on the channel output ($\mathcal{E}(\rho_0)$) that results in an outcome related to $M$ with probability

$$p = \text{Tr}(\mathcal{E}(\rho_0)M) = \frac{m_0}{\sqrt{n}} + \sum_{i=1}^{n-1} \lambda_i m_i \theta_i.$$ 

(14)

This shows that we can choose $\rho_0$ and $M$ as elements of the measurement configuration to achieve an optimal estimate.

Let us suppose that we have $N$ identical copies of the input state. Having performed $N$ measurement on the output, we compute a relative frequency ($v$) of measuring $M$, which will be close to $p$ in (14), and we can obtain an estimator $\hat{\lambda}_i$ that satisfies

$$v = \frac{m_0}{\sqrt{n}} + \sum_{i=1}^{n-1} \hat{\lambda}_i m_i \theta_i.$$ 

(15)

With $N_i$ being unknown channel parameters we need $N_i$ input state–measurement pairs, and so we will have $N_i$ equations of type (15), which can be solved uniquely.

The efficiency of estimation is usually characterized by the covariance matrix

$$\text{Var}(\hat{\lambda}_{i,j}) = E(\hat{\lambda}_{i,j})^2 - E(\hat{\lambda}_{i,j})E(\hat{\lambda}_{j,i}).$$ 

(16)

Another commonly used quantity for this is the Fisher information [14]:

$$F(\lambda_{i,j}) = \sum_a \left[ \frac{1}{p_a} \frac{\partial p_a}{\partial \lambda_i} \frac{\partial p_a}{\partial \lambda_j} \right].$$ 

(17)

Using the Cramer–Rao inequality,

$$\text{Var}(\hat{\lambda}) \geq F(\lambda)^{-1},$$

(18)

we can see that either maximizing Fisher information or minimizing the variance can lead us to optimal parameter estimation.
3. The qubit case with known channel directions

In this section we assume that the channel directions are known, and they are the Pauli directions, i.e. we have equation (10) with \( v_i = \sigma_i (i = 1, 2, 3) \). The aim is to find an optimal measurement configuration that consists of the input state(s) and POVM that maximize the Fisher information of the estimated Pauli-channel parameters.

Let us suppose that \( 1 \geq |\lambda_1| \geq |\lambda_2| \geq |\lambda_3| \) and that the POVM related to \( M \) forms a von Neumann measurement, i.e. we have

\[
M = \frac{1}{2} \left( I + \sum_{i=1}^{3} m_i \sigma_i \right),
\]

with condition \( m_1^2 + m_2^2 + m_3^2 = 1 \).

We can calculate now the Fisher information matrix from (17):

\[
F(\lambda)_{i,j} = \frac{m_i \beta_i m_j \beta_j}{1 - \left( \sum_{i=1}^{3} \lambda_i m_i \beta_i \right)^2},
\]

We have three different parameters to estimate, so to have a general setting, let us suppose that we have three different measurements represented by the Bloch vectors \( (m_1^{(1)}, m_2^{(1)}, m_3^{(1)}) \), \( (m_1^{(2)}, m_2^{(2)}, m_3^{(2)}) \), with three different input states \( (\theta^{(1)}, \theta^{(2)}, \theta^{(3)}) \). If we calculate the Fisher information matrix for each measurement \( F^{(j)} \), then we obtain the total information gained, \( F^2 = F^{(1)} + F^{(2)} + F^{(3)} \), by summing the three terms. We maximize the determinant of \( F^2 \), utilizing a straightforward extension of the Cramer–Rao inequality:

\[
\det(\text{Var}(\hat{\lambda})) \geq \det(F^2(\lambda)^{-1}) = \det(F^2(\lambda))^{-1},
\]

The above form means that the volume of the covariance matrix will be minimal, which implies that we obtain a good estimate of all \( \lambda_i \) parameters.

Using the notations \( c_i^{(j)} = m_i^{(j)} \theta_i^{(j)} \) and \( C_i = c_i^{(j)} \), the following form of the determinant is obtained:

\[
\det(F^2(\lambda)) = \det(C)^2 \prod_{j=1}^{3} \left[ 1 - \left( \lambda_1 c_1^{(j)} + \lambda_2 c_2^{(j)} + \lambda_3 c_3^{(j)} \right)^2 \right].
\]

We should maximize this determinant as a function of \( c^{(j)} \) with the proper conditions: \( |c_1^{(j)}| + |c_2^{(j)}| + |c_3^{(j)}| \leq 1 \).

**Theorem 1.** If we maximize the determinant of the Fisher information matrix with the above general measurement settings, then the optimal measurement configuration is

\[
\theta^{(1)} = (\pm 1, 0, 0), \quad m^{(1)} = (\pm 1, 0, 0),
\]

\[
\theta^{(2)} = (0, \pm 1, 0), \quad m^{(2)} = (0, \pm 1, 0),
\]

\[
\theta^{(3)} = (0, 0, \pm 1), \quad m^{(3)} = (0, 0, \pm 1).
\]

The proof is based on a modified gradient method (Mathematica [18] built-in optimizer). We choose the initial points as the elements on a grid of the state space. The result is robust; we obtain the same optimum starting from every input state–measurement pair.

Note that the same result was obtained in [1] using convex maximization arguments. This approach does not seem to be well generalizable for the multidimensional case, since even in two dimensions we have analytical problems.
If we calculate the trace of Fisher information, then the Cramer–Rao inequality implies
\[ \text{Tr}(\text{Var}(\hat{\lambda})) \geq \text{Tr}(F^\Sigma(\lambda)^{-1}) \neq \text{Tr}(F^\Sigma(\lambda))^{-1}. \]
The latter equality does not hold, so we should not maximize \( \text{Tr}(F^\Sigma) \); instead let us check which input state–measurement pair maximizes the trace of Fisher-information matrix (20).

**Theorem 2.** If the trace of the Fisher-information matrix is maximized, then the optimal input state is \( \theta = (\pm 1, 0, 0) \) and the optimal measurement is \( m = (\pm 1, 0, 0) \), with \( 1 \geq |\lambda_1| \geq |\lambda_2| \geq |\lambda_3|. \)

**Proof.** We need to maximize
\[ \text{Tr} F(\lambda) = \frac{\sum_{i=1}^{3} m_i^2 \theta_i^2}{1 - \left( \sum_{i=1}^{3} \lambda_i m_i \theta_i \right)^2}. \] (23)
From the Cauchy–Schwarz inequality we have
\[ \left| \sum_{i=1}^{3} \lambda_i \theta_i \cdot m_i \right|^2 \leq \left( \sum_{i=1}^{3} \lambda_i^2 \theta_i^2 \right) \left( \sum_{i=1}^{3} m_i^2 \right) \leq \lambda_1^2 \left( \sum_{i=1}^{3} \theta_i^2 \right) \left( \sum_{i=1}^{3} m_i^2 \right) \leq \lambda_1^2. \]
and similarly
\[ \sum_{i=1}^{3} m_i^2 \theta_i^2 \leq \left( \sum_{i=1}^{3} \theta_i^2 \right) \left( \sum_{i=1}^{3} m_i^2 \right) \leq 1. \]
The inequalities are equalities with the given input state–measurement pairs, which implies the optimality of them. \( \square \)

It is important to note that in this case one obtains information about \( \lambda_1 \) only; therefore, additional measurement configurations are needed for estimating the remaining two parameters. Using a greedy algorithm we can obtain the same result as in theorem 1.

**Corollary 1.** If we want to measure all channel parameters, then the optimal setting is that we first measure \( \lambda_1 \) in the \( \sigma_1 \) direction, then \( \lambda_2 \) in the \( \sigma_2 \) direction and finally \( \lambda_3 \) in the \( \sigma_3 \) direction.

**Proof.** Theorem 2 ensures that the first measurement configuration should be \( \theta = (\pm 1, 0, 0), m = (\pm 1, 0, 0) \). In order to obtain information about the other channel directions, we should measure the quasi-orthogonal subspace to obtain the highest possible information independently of the first measurement. This implies that \( m_1 = \theta_1 = 0 \). We should note that the proof of theorem 2 does not depend on the actual length of the sums, so we have the same optimization problem except that there are fewer terms starting from \( i = 2 \). The proof goes in the same way for that too; hence, the second optimal measurement setting is \( \theta = (0, \pm 1, 0), m = (0, \pm 1, 0), \) etc. \( \square \)

In the latest scenario let us maximize the Fisher information for \( \lambda_j \).

**Theorem 3.** The maximal value of \( F_{j,j} \) is taken when the optimal input state and the optimal measurement are both in the \( \sigma_j \) direction.

**Proof.** We have to maximize the following expression:
\[ F_{j,j} = \frac{c_j^2}{1 - \left( \sum_{i=1}^{3} \lambda_i c_i \right)^2} \to \text{max}. \] (24)
Let us suppose that \( j = 3 \), then we can write
\[
|\lambda_1 c_1 + \lambda_2 c_2 + \lambda_3 c_3| \leq |\lambda_1| |c_1| + |\lambda_2| |c_2| + |\lambda_3| |c_3| \leq |c_1| + |c_2| + |c_3| \leq 1,
\]
resulted in the same optimal configuration for qubit Pauli channels; therefore, we will use here the latest and simplest method presented in theorem 3 for obtaining the most accurate estimation of \( \lambda_j \), so we will suppose
\[
|c_1| + |c_2| + |c_3| - |c_3| + |\lambda_3||c_3| \leq 1 - (1 - |\lambda_3||c_3|).
\]
We obtain that
\[
F_{3,3} \leq I(|c_3|) = \frac{|c_3|^2}{1 - (1 - |\lambda_3||c_3|)^2}.
\]
Simple calculation gives that \( \partial I(|c_3|)/\partial |c_3| \geq 0 \), so \( I(|c_3|) \) is maximal if \( |c_3| = 1 \), and we have \( F_{3,3} \leq \frac{1}{1 - |\lambda_3|^2} \), and the equation is true for the input state and measurement given in the statement.

If we use the optimal input state–measurement pairs for each \( j \), we get here the same result as in the previous cases.

4. Multidimensional case

In this section our aim is to obtain a result similar to section 3 in higher dimension. First we note that all the three cases of maximizing measures of the Fisher information matrix (determinant, trace, independently) resulted in the same optimal configuration for qubit Pauli channels; therefore, we will use here the latest and simplest method presented in theorem 3 for obtaining the most accurate estimation of \( \lambda_j \), Thus, in the multidimensional case, we want to find the optimal input state (5) and measurement (7) combination for which \( F_{j,j} \) is maximized.

4.1. The maximal value of the Fisher information

We can calculate the Fisher-information matrix from (14) directly
\[
F_{j,j} = \frac{c_j^2}{\frac{m_0}{|\lambda_j|^2} + \sum \lambda_i c_i} \left( 1 - \frac{m_0}{|\lambda_j|^2} - \sum \lambda_i c_i \right)
\]
using \( c_j = \sum_{j \in A_j} \theta_j m_j \), where we summarize the elements corresponding to \( \lambda_i \).

It is intuitively clear that if we want to estimate \( \lambda_j \) we have to measure in the direction of \( \lambda_j \), so we will suppose \( m \in A_j \) and \( \theta \in A_j \).

Then \( c_j = 0 \), if \( i \neq j \), so we obtain
\[
I(m_0, d) := F_{j,j} = \frac{d^2}{\left( \frac{m_0}{|\lambda_j|^2} + \lambda_j d \right) \left( 1 - \frac{m_0}{|\lambda_j|^2} - \lambda_j d \right)}.
\]
where \( d \) is defined in (9), and now \( d = \sum c_i = c_j \). Calculating \( \partial I(m_0, d)/\partial d \) we observe that the Fisher information decreases for negative \( d \) values, and increases for positive \( d \) values, so it takes its maximum at the minimal or at the maximal value of \( d \). Using the definition of \( d \) we can obtain a general upper and lower bounds:
\[
0 \leq p = \frac{m_0}{\sqrt{\pi}^2} + d \leq 1 \implies d \in \left[ -\frac{m_0}{\sqrt{\pi}}, 1 - \frac{m_0}{\sqrt{\pi}} \right].
\]
We can assume that \( m_0 \leq \sqrt{n}/2 \), since if \( m_0 > \sqrt{n}/2 \) then we can use \( I - M \) instead of \( M \). If we substitute the lower bound \( d = -m_0/\sqrt{\pi} \) into (26), we see that it is an increasing function of \( m_0 \), so it is maximal if \( m_0 = \sqrt{n}/2 \). Using the maximal value of \( d = (1 - m_0)/\sqrt{\pi} \), we have the same value for \( F_{j,j} \) when \( m_0 = \sqrt{n}/2 \). Therefore, the Fisher information is maximal if the value of \( d \) is maximal.
We can use the Cauchy–Schwarz inequality to construct another estimation

\[ |d + m_0 \theta_0| = \left| \sum_{i=0}^{n^2-1} m_i \theta_i \right| \leq \sqrt{\sum_{i=0}^{n^2-1} m_i^2} \sqrt{\sum_{i=0}^{n^2-1} \theta_i^2}. \]  

(28)

Furthermore, we have the relationship

\[ \sum_{i,j} \theta_i^2 = \sum_i \theta_i \text{Tr}(v_i \cdot v_j) = \text{Tr} \left( \sum_i \theta_i v_i \cdot \sum_j \theta_j v_j \right) = \text{Tr} \rho^2 = \sum_{k=1}^{n} \mu_k^2, \]  

(29)

where \( \mu_k \) are the eigenvalues of \( \rho \). In addition, we know that \( \mu_k \geq 0 \) and \( \sum_{k=1}^{n} \mu_k = 1 \).

Assume that the eigenvalues have multiplicity \( K \), then it is easy to obtain that

\[ \sum_{k=1}^{n} \mu_k^2 \leq \frac{1}{K}. \]  

(30)

We can make a similar upper bound for \( m_i \) in (28); thus, we obtain the general upper bound

\[ |d + m_0 \sqrt{n}| \leq \frac{m_0 \sqrt{n}}{K} \implies d \leq -\frac{m_0}{\sqrt{n}} + \frac{m_0 \sqrt{n}}{K}. \]  

(31)

It is easy to check that the Fisher information increases in \( m_0 \) if \( d = \frac{m_0 \sqrt{n}}{K} - \frac{m_0}{\sqrt{n}} \) (upper bound from (31)), while the Fisher information decreases in \( m_0 \) if \( d = 1 - \frac{m_0}{\sqrt{n}} \) (upper bound from (27)). Therefore, the Fisher information is maximal if the two upper bounds are equal; hence \( m_0 = \frac{K}{\sqrt{n}} \) and \( d = \frac{1}{n-K} \), and we can conclude the following statement.

**Theorem 4.** If \( m \in A_j \) and \( \theta \in A_j \) and quantum states in \( A_j \) have multiplicity \( K \), then the maximal achievable Fisher information is

\[ I_{\text{max}} = \frac{1}{(1 - \lambda_j) \left( \frac{K}{n-K} + \lambda_j \right)}. \]  

(32)

We can use the above result to obtain some general statements.

**Corollary 2.** If \( A_j \) is isomorphic to \( \mathbb{C}^k (k \geq 2) \), then the maximal Fisher information is

\[ I_{\text{max}}^{(M_k, \mathbb{C}^k)} = \frac{1}{(1 - \lambda_j) \left( \frac{1}{m-k} + \lambda_j \right)}. \]

**Proof.** If \( A_j \) is isomorphic to \( \mathbb{C}^k \) then the eigenvalues are all with multiplicity \( K = n/k \), and theorem 4 implies the result.

**Corollary 3.** If \( A_j \) is isomorphic to \( M_m (m \geq 2) \), then the maximal Fisher information is

\[ I_{\text{max}}^{(M_k, M_m)} = \frac{1}{(1 - \lambda_j) \left( \frac{1}{m-1} + \lambda_j \right)}. \]

**Proof.** If \( A_j \) is isomorphic to \( M_m \), then the eigenvalues are all with multiplicity \( K = n/m \), and theorem 4 implies the result.
Example 3. Using the decomposition in example 1 we can obtain the maximal Fisher information:

\[ I_{\text{max}}^{(M_4:C_2)} = \frac{1}{(1 - \lambda_j)(1 + \lambda_j)}, \quad \forall j. \]

Example 4. Using the decomposition in example 2 we can obtain the maximal Fisher information:

\[ I_{\text{max}}^{(M_4:M_2)} = \frac{1}{(1 - \lambda_j)(1 + \lambda_j)}, \quad \text{for } j \in \{1, 2\} \]

and

\[ I_{\text{max}}^{(M_4:C_4)} = \frac{1}{(1 - \lambda_j)(\frac{1}{2} + \lambda_j)}, \quad \text{for } j \in \{3, 4, 5\}. \]

We can see that the maximal Fisher information strongly depends on the algebraic structure of \( A_j \) even in this simple scenario, and for the same \( \lambda_j \) values we can obtain more information if \( A_j \) is isomorphic to \( C_4 \) than if \( A_j \) is isomorphic to \( M_2 \) or \( C_2 \), since \( I_{\text{max}}^{(M_4:C_4)} > I_{\text{max}}^{(M_4:M_2)} > I_{\text{max}}^{(M_4:C_2)} \).

4.2. The optimal measurement configuration

It is in general hard to give an optimal input state–measurement pair, since it will depend on the structure of \( A_j \), but we can characterize them.

To obtain the upper bound (31), we used two inequalities, so if we want to find the optimal measurement settings, both inequalities should be satisfied with an equality. In (28) the Cauchy–Schwarz inequality is sharp if \( m \) and \( \theta \) are multiples of each other. On the other hand, the maximum is achieved in (30) when we have \( \mu_k = \frac{1}{K} \) for \( K \) different \( k \) value, and \( \mu_k = 0 \) in the other cases. From this and from \( m_0 = \frac{K}{\sqrt{m}} = K \cdot \theta_0 \) follows that the optimal measurement \( (m_{\text{opt}}) \) is an arbitrary von Neumann measurement in \( A_j \) with the minimal rank \( (K) \), and the optimal input state is \( \theta_{\text{opt}} = \frac{m_{\text{opt}}}{\theta} \).

Let us also note that the proof works either for positive or negative \( \lambda_j \); however, there are some asymmetries: the maximum value is \( \lambda_j \leq 1 \), but from (12) we obtain \( \lambda_j \geq -\frac{1}{n-1} \).

5. Unknown channel directions in the qubit case

In the previous sections we obtained results for known, given channel directions. In this section we propose a method to estimate the direction of axes too \((\varphi_i)\), besides the magnitude of contraction \((\lambda_i)\). Obtaining the optimal measurement scenario is more difficult here; therefore, the efficiency of the proposed algorithm is only investigated using simulations.

5.1. The estimation scheme

In the qubit case, the quantum states can be described with Bloch vectors (2); hence, they can be considered as elements of the \( \mathbb{R}^3 \) space. Therefore, the effect of the Pauli channel can be described with a linear operation on this space in the following way:

\[ \varepsilon (\rho_0) = \rho_{0'}, \quad \text{iff} \quad \theta^* = A \theta, \]

where \( A \) denotes the appropriate linear operator.

The channel has six real parameters: three angles \((\varphi_1, \varphi_2, \varphi_3)\) and three eigenvalues \((\lambda_1, \lambda_2, \lambda_3)\); the angles determine the channel directions, and the eigenvalues are the parameters of the contraction. We will use the following decomposition of \( A \):

\[ A(\lambda_1, \lambda_2, \lambda_3, \varphi_1, \varphi_2, \varphi_3) = R_z R_x A R_z^{-1} R_y^{-1} R_z^{-1}. \]
where
\[
R_z(\phi_1) = \begin{pmatrix}
\cos \phi_1 & -\sin \phi_1 & 0 \\
\sin \phi_1 & \cos \phi_1 & 0 \\
0 & 0 & 1
\end{pmatrix}, \\
R_y(\phi_2) = \begin{pmatrix}
\cos \phi_2 & 0 & -\sin \phi_2 \\
0 & 1 & 0 \\
\sin \phi_2 & 0 & \cos \phi_2
\end{pmatrix}, \\
R_x(\phi_3) = \begin{pmatrix}
1 & 0 & 0 \\
0 & \cos \phi_3 & -\sin \phi_3 \\
0 & \sin \phi_3 & \cos \phi_3
\end{pmatrix}, \\
\Lambda = \begin{pmatrix}
\lambda_1 & 0 & 0 \\
0 & \lambda_2 & 0 \\
0 & 0 & \lambda_3
\end{pmatrix}.
\]

We choose three input qubits, such that their Bloch vectors \((\theta^{(1)}, \theta^{(2)}, \theta^{(3)})\) are linearly independent. Using notation \(\theta = [\theta^{(1)}, \theta^{(2)}, \theta^{(3)}] \in \mathcal{M}_3(\mathbb{R})\), we have the following equation for output qubits:
\[
\theta^* = A \cdot \theta. \tag{35}
\]

Thereafter, we estimate the output qubits of the channel with three projective measurements: \((M^{(i)}, I - M^{(i)}), \quad i = 1, 2, 3\), where \(M^{(i)} = \frac{1 + m^{(i)}}{2}\). If we denote the probability of the outcome of measuring \(M^{(i)}\) for the input state \(\theta^{(j)}\) by \(P_{m^{(i)}, j}\), and using the notation \(\bar{M} = [m^{(1)}, m^{(2)}, m^{(3)}]\), we obtain
\[
P = \frac{1 + M \cdot \theta^*}{2}. \tag{36}
\]

Let \(N_{i,j}\) denote the number of measurements corresponding to \(m^{(i)}\) and \(\theta^{(j)}\), and \(N_{i,j}(\rightarrow)\) the number of \(M^{(i)}\) outcomes, then the relative frequency \(v_{i,j} = \frac{N_{i,j}(\rightarrow)}{N_{i,j}}\) will be an unbiased estimator for \(P_{m^{(i)}, j}\). Then using (36) we obtain an estimation on \(\theta^*\):
\[
\hat{\theta}^* = \left(\bar{M}^T\right)^{-1} (2v - 1), \tag{37}
\]

where \(v\) is the matrix with elements \(v_{i,j}\), and since it is linear in \(v\) we will also obtain an unbiased estimation for \(\theta^*\).

From (35) we can estimate the channel effect in the form
\[
\hat{A} = \hat{\theta}^* \cdot \hat{\theta}^{-1}. \tag{38}
\]

The estimated channel effect determines the parameter estimations (34), so we can construct equations for determining the channel parameters in the form
\[
A(\hat{\lambda}_1, \hat{\lambda}_2, \hat{\lambda}_3, \hat{\phi}_1, \hat{\phi}_2, \hat{\phi}_3) = \frac{\hat{A} + \hat{A}^T}{2}. \tag{39}
\]

We use symmetrization of \(\hat{A}\) since we know that \(A\) is symmetric. Therefore, it is enough to use only the upper triangular part of \(A\) for parameter estimation, because there are only six parameters.

We estimate the channel parameters by solving matrix equation (39).

5.2. The efficiency of the estimation

The parameters of the Pauli channel are fixed, and we can consider the input qubits and measurements on the output qubits as optimization variables. Our aim is to find the optimal measurement setup by solving the minimization problem:
\[
\mathcal{V} = (1 - c) \left( \sum_{i=1}^{3} \text{Var}(\hat{\lambda}_i) \right) + c \left( \sum_{i=1}^{3} \text{Var}(\hat{\phi}_i) \right) \rightarrow \text{min}. \tag{40}
\]
A constant $c \in [0, 1]$ is used here to express our preference with respect to the two sets of parameters. With $c = 1$ we can minimize the estimation error of the axis direction, and with $c = 0$ we can minimize the error of the contraction magnitudes, while by using $0 < c < 1$ we can minimize them simultaneously.

The problem is that the equations in (39) are nonlinear in angle parameters, so the estimators will not be unbiased even if we know that $\hat{A}$ is unbiased. Furthermore, the variances in (40) cannot be expressed analytically. Therefore, one may investigate the solution of the above optimization problem by using simulations. In this case, we generate $K$ realizations of the random variable $\nu$ for a given $M$ and $\theta$, using the built-in random function of Mathematica [18]. From each realization of $\nu$ we obtain an estimate for $\hat{A}$, and by solving equation (39) we can estimate the channel parameters, and from that we can obtain the empirical mean square error (MSE):

$$\text{MSE}(\hat{\theta}) = \frac{1}{K} \sum_{k=1}^{K} (\hat{\theta}_k - \theta)^2,$$

(41)

where $\theta$ is one of the estimated parameters ($\theta = \lambda_i \phi_i$) and $\hat{\theta}_k$ is the solution that is obtained from (39) for the $k$th realization. Note that in (41) the MSE is used instead of the variance. We know that

$$\text{MSE}(\hat{\theta}) = \text{Var}(\hat{\theta}) + (\mathbb{E}(\hat{\theta}) - \theta)^2,$$

(42)

so aside from showing that the variance vanishes, the MSE tending to zero indicates that the estimation method is asymptotically unbiased.

By substituting the quantities (41) into (40) instead of the variances, we obtain an estimate $\hat{\nu}$, which is a good approximation of $\nu$ if $K$ is large enough ($\text{MSE}(\hat{\theta}) \rightarrow \text{MSE}(\hat{\theta})$) and the bias of $\hat{\theta}_k$ is small (which turns out to be true for large $N_{ij}$).

5.3. Simulation results

For the simulation investigations, the following constant system parameters were used:

$\lambda_1 = 3/4, \lambda_2 = 1/2, \lambda_3 = 1/4, \phi_i = 0$, for $i \in \{1, 2, 3\}$.

In each direction the same number of measurements $N_{ij} = N$ were used, and we repeated the whole measurement process many times ($K = 1000$) to ensure that the values of $\hat{\nu}$ are close to the real MSE.

First we remark that $M$ and $\theta$ are invertible, and we used their inverses in our calculations. Moreover, the results with known channel directions reported in the earlier sections suggest that they are probably orthogonal in the optimal case. Therefore, $\hat{\nu}$ was investigated as a function of the input state triplets with a randomly chosen but fixed $M$ measurement matrix. For this purpose, some pure states were used as input states (the angle between them was gradually increased), and the value $\text{det}(\theta)$ was used as a measure of orthogonality. The result of this investigation can be seen in figure 1(a), where it is visible that a smaller variance was obtained for the orthogonal case. The same result was observed when a fixed input state triplet was chosen and the measurement directions were changed: the orthogonal measurement direction case produces the smallest variance. As a conclusion, orthogonal matrices for $M$ and $\theta$ were found to be optimal.

Thereafter, we investigated the variance for three different $c$ values: $c = 0$, $c = 0.5$ and $c = 1$. For this purpose, a fixed random orthogonal measurement direction set, and input state triplet $\hat{\theta} = I$ were used. The result is shown in figure 1(b). It is seen that $N \cdot \hat{\nu}$ is nearly constant for all $\hat{c}$ values (except for low $N$ values), so the MSE converges to 0 in $1/N$ order. This means
that we can obtain accurate state estimation for all cases if $N$ is large enough. The difference is in the coefficients: if $c = 0$, then the variance is smaller, so we can estimate $\lambda$ more accurately than the angles from the same number of measurements. This is not surprising, since then the estimators are linear.

A further useful observation is that for every $0 < c < 1$ the variance can be expressed as a superposition of the $c = 0$ and $c = 1$ cases, so we can separate the problem into two parts: to measuring the directions of the channel and to measuring the magnitude of contractions. The latter has been discussed in section 3 and an analytical solution was found, so it is enough to examine the $c = 1$ case by using simulation.

Section 3 would suggest that the optimal input states and measurements should be in the channel directions. However, this is not true in the $c = 1$ case: by rotating $M$ we could achieve a better value for $\hat{V}$ than with $M = \theta = I$. Therefore, further investigations are needed to find the optimal input states and measurements in the $c = 1$ case, i.e. when we want to estimate the channel directions.

6. Discussion and conclusions

In this paper we examined different scenarios to find optimal parameter estimation schemes for Pauli channels that consist of the optimal input state and the optimal two-element POVM.

Assuming known channel directions, both the qubit and the $n$-dimensional cases were discussed in detail. In the two-dimensional case we have shown that the optimal measurement–input state pairs should be directed towards the channel directions. We have achieved the same optimal result in three different ways: first we gave a jointly optimal estimation for the parameters, then we maximized the total achievable information in each step and finally we constructed the best estimation for $\lambda_j$ independently.

In the $n$-dimensional case the independent $\lambda_j$ estimation method was used because of its simplicity. We have obtained similar results that in the qubit case: in the optimal estimation scheme of $\lambda_j$, the measurement and input states are in the corresponding subalgebra $A_j$, and they are multiples of each other. Furthermore, the measurement should be a von Neumann measurement with a minimal rank in $A_j$. This implies that the optimal Fisher information depends on the algebraic structure of $A_j$.

Finally, we investigated the case when we have to estimate the channel directions, too. The quantity (40) was used to measure the quality of estimation, and the estimates were obtained.
by its minimization. A tuning parameter was introduced to weight the variance of channel direction estimations to that of contractions. It was found that to analytically determine the optimal estimation scheme of channel directions is hard even in the qubit case. Our simulation investigations show that the MSE of the proposed method converges to zero with the order of $1/N$, where $N$ is the number of used measurements, so we can achieve the same estimation order as in $\lambda_j$ case in spite of nonlinearity. Our simulations also indicate that we do not necessarily achieve the smallest variance if we measure in the channel direction.

Acknowledgment

This research was supported in part by the Hungarian Research Fund through grant K83440. This project was supported by the New Széchenyi Plan (project ID: TÁMOP-4.2.1/B-09/1/KMR-2010-0002).

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