CONSTRAINTS ON LEMAÎTRE–TOLMAN–BONDI MODELS FROM OBSERVATIONAL HUBBLE PARAMETER DATA

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ABSTRACT

We use the observational Hubble parameter data (OHD), both the latest observational data set (Stern et al.; referred to as SJVKS) and the simulated data sets, to constrain Lemaître–Tolman–Bondi (LTB) void models. The necessity of the consistency check on OHD itself in the LTB cosmology is stressed. Three voids are chosen as test models and are constrained using the Union2 data set of SN Ia as well as OHD. Despite their different parameterization, the results from our test models show some similarities; e.g., the best-fit voids obtained from OHD are all considerably broader than those from SN Ia. Due to the small size of the SJVKS data set, the constraints are not conclusive. The constraining power of the future OHD observations are therefore investigated through a figure of merit (FoM) analysis based on the Monte Carlo simulated data. We found that, in the case that the future OHD become more consistent with SN Ia, the results from the test models are almost unanimous: (1) as many as 32 OHD data points at the SJVKS-like uncertainty level are needed to give a higher FoM than the Union2 data set of SN Ia, (2) precise observation helps reduce this required number, (3) increasing the survey depth does not always increase the FoM. On the other hand, if the future OHD and the Union2 data set keep favoring different voids, in a similar manner as they do at present, the 1σ confidence regions obtained from the two probes should finally separate. We test this conjecture and find that the minimum observational requirement (the size of the data set, the uncertainty level, and the survey depth) for this inconsistency to emerge depends strongly on the void model.

Key words: cosmology: miscellaneous – methods: numerical

1. INTRODUCTION

Since the direct supernovae measurement of the deceleration parameter (Riess et al. 1998; Perlmutter et al. 1999), the last decade has seen a solidification of the standard cosmological model, which has about 75% dark energy driving the acceleration of a flat Friedman–Robertson–Walker (FRW) universe. A very long array of observations, including cosmic microwave background (CMB; Komatsu et al. 2011), baryon acoustic oscillations (BAO; Percival et al. 2010), weak lensing (Schrabback et al. 2010), etc., are consistent with this model. Yet a radically different scenario has attracted considerable attention recently as an alternative model choice to fit the observations. The idea is to suppose we are very near to the center of a large spherical void, which is itself embedded in an otherwise homogeneous background (Célerrier 2000). The dynamics of this kind of model can be described in the Lemaître–Tolman–Bondi (LTB) metric, hence the term LTB models. Constraints from supernovae, CMB, BAO, and H0 have shown that this model cannot be ruled out (García-Bellido & Hu 2008; Biswas et al. 2010; but see Moss et al. 2011; Zibin et al. 2008). As the cross-check from the consistency of the results based on independent evidences is an effective, if not the only, way to pin down the correct cosmological model (Peebles & Ratra 2003), more observations are needed to better discriminate the standard FRW and the LTB scenarios.

Observational Hubble parameter data (OHD) have been measured through the aging of passively evolving galaxies (Jimenez & Loeb 2002; Simon et al. 2005) and the BAO effect (Gaztañaga et al. 2009). Recently, a third method has been proposed to obtain OHD from the gravitational wave standard sirens (Nishizawa et al. 2011). OHD is shown to be consistent with the standard cosmological model (Lin et al. 2009; Stern et al. 2010a). Various other FRW models have also been tested against OHD (Samushia & Ratra 2006; Yi & Zhang 2007; Zhang & Zhu 2008; Zhai et al. 2010; Xu & Wang 2010; see Zhang et al. 2010 for; a recent review). The potential of future OHD observations in constraining FRW models is explored via simulation in Ma & Zhang (2011). Only recently this data set was used to test LTB models in February et al. (2010; hereafter FLSC).

The current work deals with a purpose similar to FLSC, as we believe more discussion is needed, both on the subtleties when using OHD in LTB cosmology, and on its potential constraining power as the data set enlarges in the future. The investigation of these two aspects, as well as the current constraints on various void models, constitutes the main content of this paper.

We briefly review the LTB cosmology and introduce several test models in Section 2. In Section 3, we present the physics and assumptions involved in obtaining OHD. This analysis then guides our choice in picking the proper observable in LTB cosmology correspondent to the OHD data set and signifies the necessity for a consistency check on OHD itself in the LTB cosmology. Section 4 illustrates how, in the LTB cosmology, the likelihood analysis should be done when OHD is used; the best-fit void parameters obtained in this section are used in Section 5 to compare between the LTB void models and the standard model. The potential constraining power of the future OHD observations is explored in Section 6 through a figure of merit (FoM) analysis; in the same section we also discuss whether the future OHD data set can be used, when combined with the Union2 data set, to exclude the void models. We conclude in Section 7.

2. LTB MODELS

The LTB solution to the Einstein field equations describes the dynamics of a spherically symmetric dust universe. We take a
brief review of the basic formulas in LTB cosmology, then we introduce the test LTB models used in our analysis.

2.1. Basics of LTB Cosmology

The background LTB line element is written as

\[ ds^2 = -dt^2 + \frac{A'(r,t)^2}{1 - k(r)}dr^2 + A^2(r,t)d\Omega^2, \]  

(1)

where \( \dot{\theta} \) denotes \( \partial/\partial t \), and \( k(r) \) is an arbitrary function of \( r \), playing the role of spatial curvature. The stress-energy tensor of the mass source is given by

\[ T^\mu_\nu = -\rho(r,t)\delta^\mu_\nu, \]  

(2)

Note that the FRW metric can be recovered by imposing \( A(r,t) = a(t)r \) and \( k(r) = k_r^2 \). The Einstein field equations read

\[ H^2_\perp + 2H_\parallel H_\parallel + \frac{k(r)}{A^2} + \frac{k'(r)}{AA'} = 8\pi G\rho_m, \]  

(3)

\[ \dot{A}^2 + 2A\ddot{A} + k(r) = 0, \]  

(4)

where \( \dot{\theta} \) denotes \( \partial/\partial t \), and

\[ H_\perp \equiv \frac{\dot{A}^2}{A^2}, \]  

(5)

\[ H_\parallel \equiv \frac{\dot{A}'}{A'}, \]  

(6)

are the expansion rates at the transverse and longitudinal directions, respectively. Integrating Equation (4), we get

\[ H^2_\perp = \frac{M(r)}{A^3} - \frac{k(r)}{A^2}, \]  

(7)

where \( M(r) \) is another arbitrary function. The evolution of \( A(r,t) \) can be obtained by integrating Equation (7); the standard parametric solutions are as follows (Moss et al. 2011):

\[ A(r,t) = \frac{M(r)}{k(r)}(1 - \cosh\eta), \quad t - t_B(r) = \frac{\sinh\eta}{-k(r)^{1/2}}(\sinh\eta - \eta), \]  

(8)

for \( k < 0 \); and

\[ A(r,t) = \frac{M(r)}{k(r)}(1 - \cos\eta), \quad t - t_B(r) = \frac{\sin\eta}{k(r)^{1/2}}(\eta - \sin\eta), \]  

(9)

for \( k > 0 \); and

\[ A(r,t) = \left( \frac{9M(r)}{2} \right)^{1/3} (t - t_B(r))^{2/3}, \]  

(10)

for \( k = 0 \). Here, \( t_B \) is another arbitrary function, referred to as the “bang time,” due to the singularity behavior at \( t = t_B \).

Substituting Equation (7) into Equation (3) gives

\[ \frac{M'(r)}{AA'} = 8\pi G\rho_m. \]  

(11)

Evaluating Equation (7) at present time gives the interdependence between \( M(r) \) and \( k(r) \), which, with a notation similar to that in FRW cosmology, can be parameterized as

\[ M(r) = H^2_\perp(r)\Omega_m(r)A^3_0, \quad k(r) = H^2_\perp(r)(\Omega_m(r) - 1)A^2_0, \]  

(12)

where a subscript 0 denotes the current value (\( t = t_0 \)) of a quantity. Substituting Equation (12) and \( A(t = t_0, r) = A_0(r) \) into Equation (8) (or Equations (9) and (10), depending on the value of \( k(0) \)), one can get the age of the universe as a function of \( r \), given by

\[ t_0 - t_B(r) = \frac{1}{H_{1,0}(r)}F(\Omega_m), \]  

(13)

where

\[ F(x) = \begin{cases} \frac{-\sqrt{1 - x}\sin^{-1}\sqrt{x}}{(x - 1)^{1/2}} & x > 1 \\ \frac{2/3}{\sqrt{1 - x}\sin^{-1}\sqrt{x}} & x < 1 \end{cases}, \]  

(14)

In order to compare the theory with the observation, one needs to associate the coordinates with the redshifts. This can be done by using the redshift equation (Alnes & Amarzguioui 2006; Enqvist 2008)

\[ \frac{dz}{dt} = -\frac{\dot{A}'(r,t)}{A'(r,t)}, \]  

(15)

and the equation of \( dt \) and \( dr \) on the light cone

\[ \frac{dt}{dr} = -\frac{A'(r,t)}{\sqrt{1 - k(r)}}. \]  

(16)

As a Cauchy problem, the differential equations can be solved numerically for any redshift once the profile of \( \Omega_m \) and \( H_{1,0} \) are set. We do the equation integration with a modified open-source software easyLTB (Garcia-Bellido & Haugbolle 2008), where \( A_0(r) = r \) is chosen to fix the gauge. The calculation process for the cases when \( \Omega_r \neq 0 \) is as follows.

1. Start from \( t = t_0, r = 0 \).
2. Solve for \( \eta \) from the second equation of Equation (8) or Equation (9).
3. Obtain the expression of \( A, A' \), and \( A' \) from the first equation of Equation (8) or Equation (9), \( \eta' \) and \( \dot{\eta} \) in the result can be expressed as functions of \( A \) using the second equation.
4. Solve \( A' \) from Equations (3), (5), (6), (11), and (12).
5. Integrate Equations (15) and (16) over \( z \) to get the new \( t(z) \) and \( r(z) \).
6. Iterate Steps 2–5 till the desired \( z \) is obtained.

The original easyLTB does not support the void models with \( \Omega_k \leq 0 \). We modify the code to allow for negative \( \Omega_k \). For \( \Omega_k = 0 \), it is no longer valid to use the parametric forms. However, as both Equations (8) and (9) converge to Equation (10) when \( \Omega_k \rightarrow 0 \), in practice one can always approximate the correct answer with the parametric form by setting \( \Omega_k \) to a small nonzero value: following the calculation process above, letting \( \Omega_k = 10^{-6} \) would cause a 10\(^{-3}\) error in \( \eta \) and lead in turn to an 10\(^{-9}\) error in \( A \) for each integration step. The error will accumulate during the integration but should not be larger than 1% for any sensible number of integration steps (normally smaller than 100,000).
One can tell the angular diameter distance directly from the metric form, Equation (1), \[ d_A(z) = A(r(z), t(z)), \] hence the luminosity distance (the distance duality relation holds as long as one stays with the metric theory of gravity and does not consider any exotic physics (Etherington 1933; Bassett & Kunz 2004)), \[ d_L(z) = (1 + z)^2 A(r(z), t(z)). \]

The LTB dynamics discussed above consider only matter, so they break down at high redshifts where radiation becomes important. This is one reason why a large family of LTB models are voids that are embedded in an FRW background universe (Biswas et al. 2010; Zibin et al. 2008; Moss et al. 2011). Moreover, out of consideration for meeting the prediction of the inflation, an Einstein–de Sitter universe is often chosen as this background (Garcia-Bellido & Haugbolle 2008; FLSC). The LTB models so constructed are referred to as EdS voids. These types of voids are said to be ruled out as a whole according to Moss et al. (2011; see also Zibin et al. 2008), as those voids which fit the CMB data will give too low a $H_0$ that will contradict the local observation. Clarkson & Regis (2011) found, however, that there are ways out if one takes into account the radiation when joining the voids to the background FRW universe, for instance, by introducing more relativistic degrees of freedom, or by considering a varied baryon fraction and/or a baryon–photon ratio.

Our aim in this work is mainly to explore the constraining power of OHD data, compared with SN Ia data, both of which are low-redshift observations. Therefore, we do not have to take into account the role played by the radiation. Besides, whether the EdS voids are feasible or not, as test models they could in any case serve as a test bench of the constraining power of different observational probes. Third, as mentioned above, the EdS background is consistent with inflation, and is therefore a theoretically conservative choice. For these reasons, the test models we use in this work are all EdS voids. Also, we assume that there are no isocurvature modes, i.e., variation in the baryon fraction or in the baryon–photon ratio, correlated with the void, because these will enhance the inhomogeneity in the structure formation time which might make OHD itself invalid, not to mention using it to constrain models (see the discussions in Section 3).

2.2. Void Models

The radial inhomogeneity of LTB solutions leaves an almost totally arbitrary form of the radial profiles of $\Omega_m$, $H_{1,0}$, and $t_0$. For $\Omega_m$, the much discussed void-like profiles are favored. This is because of the requirement that the LTB model should have an expansion rate decreasing with the distance to the center, which in turn observationally resembles an expansion rate increasing with time in FRW models.

Besides, the gradients in the bang time, $t_0$, correspond to a currently non-vanishing decaying mode (Silk 1977; Zibin 2008). This would imply a very inhomogeneous early universe, and hence violate inflation. More importantly for the current work, this will lead to great inhomogeneities in the galaxy formation time and make the OHD data set invalid. So $t_0$ must be a constant, and we set it to be 0. $H_{1,0}$ for the desired void models is then given by substituting $t(r) \equiv t_0$ into Equation (13): \[ H_{1,0}(r) = H_0 F(\Omega_m), \] where $H_0 \equiv 1/t_0$, (20) and $F$ is given by Equation (14).

With the relation between $H_{1,0}$ and $\Omega_m$ determined by Equation (19), we need only fix the specific profile of $\Omega_m$. The constrained GBH (CGBH) model (Garcia-Bellido & Haugbolle 2008) is chosen to be one of our test models, \[ \Omega_m(r) = 1 + (\Omega_0 - 1) \left[ \frac{1 - \tanh[(r - r_0)/2\Delta r]}{1 + \tanh[r_0/2\Delta r]} \right], \] where $\Omega_0$ describes the density at the symmetric center, $r_0$ is the characteristic size of the void, and $\Delta r$ describes the steepness of the void near the edge. Figure 1 illustrates profiles of $\Omega_m$ and $H_{1,0}$ in the CGBH model, with $H_0$, $r_0$, and $\Omega_0$ fixed at 74 km s$^{-1}$ Mpc$^{-1}$, 0.05, and 2 Gpc, respectively. Different colors (red, blue, and brown) stand for different values of $\Delta r/r_0$ (0.1, 0.3, and 0.9, respectively). It can be seen that a small $\Delta r/r_0$ corresponds to an $\Omega_m$ profile that is flat at the origin, i.e., $\Omega_m|_{r=0} = 0$, while large $\Delta r/r_0$ more pointed ones ($\Omega_m|_{r=0} \neq 0$). Because of this feature, and also the consideration that $\Delta r$ should not be greater than unity, we use $dr \equiv \Delta r/r_0$ as a free parameter instead of $\Delta r$. Note that despite the notation, $\Omega_0$ does not reflect the exact profile of the matter density. At a given cosmic time $t^*$, the density profile $\rho_m(r, t^*)$ is determined through Equation (3).

The value of $\Omega_m'$ at the origin turns out to be of great importance. Defining the curvature as $K \equiv k/r^2$, one could tell from Equations (12) and (19) that the $r$-derivative of $K$, $K'$, is directly related to $\Omega_m$. Indeed, it is straightforward to show that an LTB model with homogeneous bang time is smooth at the center, $K'|_{r=0} = 0$, if and only if $\Omega_m|_{r=0} = 0$ and $\frac{d^2}{dr^2}|_{r=0}$ is finite. These kinds of voids are referred to here as the “smooth voids.” Particularly, the CGBH void is a smooth void only when $dr \rightarrow 0$. It has been argued that the smooth voids are more natural. More importantly, the luminosity distance–redshift relation in the smooth voids differs qualitatively from that of the FRW models, which could be used to differentiate these two scenarios from the future supernovae data (Clifton et al. 2008). Keeping this result in mind, we choose for the rest of our test models to always be smooth. From the discussion above, this means we look for models with $\Omega_m$ that is differentiable at the origin. The first such model we choose has the simplest

![Figure 1. $\Omega_m$ and $H_{1,0}$ as functions of $r$, with fixed $H_0$, $\Omega_0$, $r_0$, and different $dr$. Small $dr$ is seen to lead to profiles that are flat near the origin, and large $dr$ to more pointed ones.](image)
Gaussian profile:
\[
\Omega_m(r) = 1 + (\Omega_0 - 1) \exp \left( -\frac{r^2}{2\sigma^2} \right),
\]
with the parameter set as \(\{\Omega_0, r_0, H_0\}\). The second smooth model is obtained through a cubic spline interpolation. The constraints we set, given three \(r\) values \(r_1 = 0 < r_2 < r_3\), for determining the \(\Omega_m\) profile, are that

1. \(\Omega_m(0) = \Omega_0\).
2. \(\Omega_m'(r=0) = 0\), i.e., a smooth void.
3. \(\Omega_m(r_2) = \Omega_2\).
4. \(\Omega_m(r_3) = 1\).
5. \(\Omega_m'(r=r_3) = 0\), which together with (4) makes this model an EdS void.

The parameter set for the Spline model is then \(\{\Omega_0, r_2, \Omega_2, r_3, H_0\}\). In practice, to make sure \(r_2\) is no larger than \(r_3\) we use \(r_2/r_3\) as a free parameter instead of \(r_2\).

Throughout the paper, the CGBH model is treated as our main model. Most of the calculation details will be described only for this model, although the main results for all three models will always be shown and discussed together.

3. OBSERVATIONAL HUBBLE PARAMETER FROM PASSIVELY EVOLVING GALAXIES

In this section, we first give a brief review on how OHD is obtained and utilized in the FRW scenario, then we will figure out the complexities that arise when moving to the LTB scenario. As mentioned in the introduction, there are currently two sources of OHD data points, those from the passively evolving galaxies and those from the large-scale structure (BAO). The latter method depends on the detailed evolution of perturbations which is not well understood in the LTB cosmology, although progress has been made (Zibin 2008; Clarkson et al. 2009). Therefore, the OHD discussed in this work refer exclusively to the former.

3.1. FRW Scenario

The Hubble parameter in the FRW models is given by
\[
H \equiv \frac{\dot{a}}{a} = -\frac{1}{1+z} \frac{dz}{dt_{\text{ct}}},
\]
where \(a\) is the scale parameter and \(dt_{\text{ct}}\) is the variation of the cosmic time due to a small change in the redshift \(dz\). One can measure \(H\) directly through the differential age method (Jimenez & Loeb 2002). We briefly review this idea below.

For any galaxy, one has
\[
T_{\text{CA}}(z) = T_F + T_{\text{GA}}(z),
\]
which simply states that the cosmic age \(T_{\text{CA}}\) at redshift \(z\) equates the summation of the time for the galaxy to form, \(T_F\), and the age of this galaxy, \(T_{\text{GA}}\). The latter can be determined spectroscopically. If we could find a group of galaxies that share a uniform formation time, i.e., \(T_F = \text{const.}\), we would then get a handle on \(dt_{\text{ct}}\) by simply measuring the age difference of those galaxies:
\[
dt_{\text{ct}}(z) = dT_{\text{CA}} = dT_F + dT_{\text{GA}}(z) = dT_{\text{GA}}(z).
\]

The passively evolving galaxies can be identified by figuring out at every redshift for the oldest galaxies which define the “red envelope.” In this case, one assumes the oldest galaxies formed at the same time, which is of course a natural assumption in an FRW universe (one may call it the galaxy formation version of the cosmic Copernican principle). By observing 24 massive galaxy clusters, determining the red envelope and differentiating the ages of the galaxies on the red envelope, Stern et al. (2010a) add two new points to those of Simon et al. (2005). The resulting data set (hereafter the SJVKS data set) is listed in Table 1.

| \(z\) | \(H(z)\) (km s\(^{-1}\) Mpc\(^{-1}\)) |
|------|---------------------|
| 0.1  | 69 ± 12             |
| 0.17 | 83 ± 8              |
| 0.27 | 77 ± 14             |
| 0.4  | 95 ± 17             |
| 0.48 | 97 ± 62             |
| 0.88 | 90 ± 40             |
| 0.9  | 117 ± 23            |
| 1.3  | 168 ± 17            |
| 1.43 | 177 ± 18            |
| 1.53 | 140 ± 14            |
| 1.75 | 202 ± 40            |

3.2. LTB Scenario

There are two rates of expansion in LTB models, defined, respectively, in Equations (5) and (6). It turns out that the longitudinal expansion rate \(H_1\) has the same form as Equation (23), hence it corresponds to the observed \(H(z)\) in Table 1. This can be seen by combining Equations (6) and (15), which gives
\[
H_1 = -\frac{1}{1+z} \frac{dz}{dt_{\text{ct}}},
\]
Ambiguities arise, however, from Equation (24) when determining \(dt_{\text{ct}}\). We have seen in the last subsection that the same formation time for the oldest galaxies is the basic assumption in obtaining OHD. However, in LTB models where the universe has a considerable background inhomogeneity, this assumption becomes unreasonable and has to be dropped. So, one question that must be answered before any data analysis is performed is whether OHD can still be regarded as a valid data set in the LTB scenario.

In answering this question, we first noticed that, albeit the overall uniform-formation-age assumption, the validity of an OHD data point requires the same formation time, only inside the redshift bin where it is obtained (Equations (23)–(25)). This is to say, OHD data points are locally defined and obtained, and therefore can be correct even when the global density, and hence the formation time of the oldest galaxies, varies greatly at different redshifts. Second, in the models’ side, a standard viewpoint (referred to as the Onion approximation) treats the LTB void universe as a group of thin shells structured together, and that inside each of these spherical shells the matter is homogeneously distributed (Biswas et al. 2007). If the bin size for determining OHD is small enough, and the background density varies only slightly inside it, we could identify these redshift bins with the homogeneous shells above and apply the Onion approximation. Third, for the SJVKS data set, the size of each bin is between 0.1 and 0.15, where the first limit is so chosen that the age evolution between the two bins is larger than the error in the age determination (Simon et al. 2005). As the precision of the age determination improves, we expect an even smaller bin size.
It is by noticing the three points above that we suppose that OHD could still be used in LTB models. In order to be sure about this, however, one needs exact knowledge about the thickness of the shell given the size of a redshift bin, as well as the steepness of the density profile at the time when the oldest galaxies formed. This information will not be learned until one obtains the constraints on the model parameters. So in practice, one has to first assume the Onion approximation holds, and check the validity of this assumption after the data fitting is done. The fitting result can be accepted only if the Onion approximation turns out to be valid. We will do this consistency check at the end of the next section.

4. CONSTRAINTS FROM THE CURRENT OHD DATA SET

4.1. Likelihood

We adopt \( \chi^2 \) statistics to determine the most likely values, as well as the confidence intervals, for the parameter sets used in our test models. Below are the main formulas using the CGBH model as an example. For a data set of \( \{ H_i^D \} \) with errors \( \{ \sigma_i \} \), \( \chi^2 \) is defined as

\[
\chi^2(H_0, \Omega_0, r_0, \Delta r) = \sum_i \frac{[H_i^D(z_i; H_0, \Omega_0, r_0, \Delta r) - H_i^{D0}]^2}{\sigma_i^2}.
\]

(27)

The likelihood function for the parameters given the data set can be obtained via Bayes’ theorem,

\[
L(H_0, \Omega_0, r_0, \Delta r|\{ H_i^D \}) = \frac{p(\{ H_i^D \}|H_0, \Omega_0, r_0, \Delta r)p(H_0, \Omega_0, r_0, \Delta r)}{p(H_0, \Omega_0, r_0, \Delta r)} \propto p(\{ H_i^D \}|H_0, \Omega_0, r_0, \Delta r) \propto \exp \left( -\frac{\chi^2}{2} \right).
\]

(28)

where the second equation comes from the fact that no prior constraints are imposed on the data and that a uniform prior is assumed for the parameters inside the respective scanning ranges. Table 2 lists the parameter range scanned in the CGBH model, for the two void models see Appendix B.

4.2. Constraints from SJVKS Data

We apply the \( \chi^2 \) analysis to the SJVKS data set of OHD, with the underlying model being our three test models. In Figure 2, we show the resulting 1\( \sigma \) and 2\( \sigma \) marginalized likelihoods for the CGBH model. As a comparison, we also plot the likelihood contours for the latest Union2 data set of SN Ia. Like in the FRW case, \( H_0 \) in LTB models is treated as a nuisance parameter for SN Ia data (see Appendix A for a discussion on this point). Table 3 lists the one-dimensional best-fit parameters of the CGBH model. Since the exact fitting result of a particular model is of only secondary importance, we put the other two models’ results in Appendix B just for reference.

Neither data set gives satisfactory constraints. Nevertheless, several instructive clues can be observed from Figure 2: (1) OHD prefers a void profile that is considerably broader than that from SN Ia; (2) both SN Ia and OHD prefer a void that is almost depleted at the origin; (3) aside from the nuisance parameter \( H_0 \), the SN Ia Union2 data set is generally more constraining than the SJVKS data set of OHD.

Note that the one-dimensional likelihood of \( dr \) from OHD is far from being Gaussian and the best-fit value seen is not so well defined. This in turn weakens the first point of the conclusions above. However, we find that the same three conclusions hold for the two smooth void models too, despite their totally different parameterization. This might be a sign that the differences are real and are model-insensitive (Figure 3). We are particularly interested in the first point: if the future data for OHD and SN Ia both hold their different preferences for the void profiles, then the smaller confidence regions they give would finally detach from each other. This would then be a sign that the void model is incorrect and should be excluded. In Section 6, we will give a rough prediction on this based on simulated OHD data.

4.3. Validity Check of the Onion Approximation

Now that we have obtained the best-fit void models, we can use them to check if the Onion approximation we have made indeed holds, i.e., if the universe can be treated as homogeneous inside the redshift bins where the OHD data points are obtained. We take the best-fit parameters given by the SJVKS data set and show in Figure 4 the density profile at different cosmic times. The inhomogeneity is seen to become more significant and the void gets deeper as the universe evolves, because of a higher expansion rate inside the void. In order to see more clearly how the density varies inside a small redshift bin with which we are more concerned, we plot in Figure 5 the relative density variation inside a \( \Delta z = 0.15 \) bin. More specifically, we calculate on a given time slice \( \Delta \rho = \rho(r,z)/\rho(r,z) \), where \( \rho(r,z) \) is the density value at \( r \), with \( r \) converted from \( z \) using the best-fit parameters. As stated in the last section, each OHD data point is obtained using galaxies that are in a redshift bin no larger than 15 bin. More specifically, we calculate on a given time slice \( \Delta \rho = \rho(r,z)/\rho(r,z) \).

3 As we do not expect a homogeneous age for all the first galaxies in the LTB models, the red envelope does not have the simple relation with the cosmic time as it does in the FRW models. But the ages of the galaxies on the red envelope can still serve as a lower bound for the ages of the universe at the corresponding redshifts.
Figure 2. Constraints on CGBH model parameters from the SJVKS data set of OHD and the Union2 data set of SN Ia. The red and brown filled regions correspond, respectively, to 1σ and 2σ confidence regions obtained from the SJVKS data set of OHD. Overplotted are the 1σ and 2σ confidence regions obtained from the Union2 data set of SN Ia, in solid and dotted lines, respectively.

Figure 3. Curvature profiles of the best-fit models from the SJVKS data set of OHD and the Union2 data set of SN Ia. OHD prefers voids with slow curvature changing near the origin, while the best-fit voids obtained from SN Ia are more cuspy.

Figure 4. Density profiles at different cosmic times. As the voids get deeper, the background density becomes more inhomogeneous.

5. COMPARISON BETWEEN BEST-FIT VOIDS AND FLAT ΛCDM

5.1. AICc Analysis

The theoretical prediction for different models at their best-fit parameters, as well as the observational data of OHD and SN Ia, are plotted in Figures 6 and 12, respectively. The current observation of these two probes are not able to exclude either. To determine which model is preferred by the data, a useful tool is the Akaike Information Criterion (AIC; Akaike 1974). The AIC for a model is

\[ \text{AIC} = 2k - 2\ln L_M, \]

where \( k \) is the number of the parameters and \( L_M \) is the maximized likelihood function. The preferred model is the one

(Wang & Zhang 2010a). We find however, that the outcome depends strongly on the models used to fit the galaxy spectra, which could sometimes yield a \( T_{GA}(0) \) greater than 14 Gyr. This number would defy all our best-fit void models, as well as the latest ΛCDM model result, 13.79 Gyr (Komatsu et al. 2011), because \( T_F \) would then be negative. So finally, we could not know for sure whether the Onion approximation is valid or not. Only qualitatively could we say that the oldest galaxies observed seem to be formed very early, so very likely the SJVKS data set is valid. More precise observations are needed to determine \( T_{GA}(0) \), hence \( T_F \). As we need to use the best-fit models to generate simulated data, we leave this as an open problem and assume from now on that the Onion approximation is justified.

We find however, that the outcome depends strongly on the models used to fit the galaxy spectra, which could sometimes yield a \( T_{GA}(0) \) greater than 14 Gyr. This number would defy all our best-fit void models, as well as the latest ΛCDM model result, 13.79 Gyr (Komatsu et al. 2011), because \( T_F \) would then be negative. So finally, we could not know for sure whether the Onion approximation is valid or not. Only qualitatively could we say that the oldest galaxies observed seem to be formed very early, so very likely the SJVKS data set is valid. More precise observations are needed to determine \( T_{GA}(0) \), hence \( T_F \). As we need to use the best-fit models to generate simulated data, we leave this as an open problem and assume from now on that the Onion approximation is justified.
so one can tell approximately the relative background density variation inside models. The dotted vertical lines indicate the data points of the SJVKS data set, \( \Delta \) of the redshift.

For OHD however, the number of the data points, \( n \), is comparable to \( k \), and the probability of overfitting becomes considerable. A second-order correction of AIC, \( \text{AIC}_c \), should be used in such cases (Sugiura 1978; Hurvich & Tsai 1989; Burnham & Anderson 2002):

\[
\text{AIC}_c = \text{AIC} + \frac{2k(2k + 1)}{n - k - 1}. \tag{30}
\]

\( \text{AIC}_c \) punishes the extra parameters more harshly, and it converges to AIC when \( n \) becomes large. Note that only the difference between the \( \text{AIC}_c \) is meaningful, so it is convenient to omit the constant in \( \ln L_M \), which is the same for the same data sets and expresses \( \text{AIC}_c \) in \( \chi^2 \):

\[
\text{AIC}_c = 2k + \chi^2_{\min} + \frac{2k(k + 1)}{n - k - 1}. \tag{31}
\]

We list the \( \text{AIC}_c \) values for our LTB test models and flat \( \Lambda \)CDM in Table 4. We discuss this result again using the CGBH model as an example. Although it has an \( \text{AIC}_c \) calculated from SN data that is fairly close to that of the flat \( \Lambda \)CDM, the CGBH model is considerably less supported by OHD. This is not to say that the CGBH model fits OHD worse than the flat \( \Lambda \)CDM. On the contrary, the minimum \( \chi^2 \) from the OHD data set are \( \chi^2_{\min}(\text{CGBH}) = 6.97 \) and \( \chi^2_{\min}(\text{\( \Lambda \)CDM}) = 7.66 \), respectively, i.e., the former fits the data better. Therefore, of the three terms in Equation (31), the third contributes the most to the difference in \( \text{AIC}_c \). As this term will get smaller as the data set enlarges, the performance of CGBH against OHD in the \( \text{AIC}_c \) test is expected to get better in the future.

### 5.2. Effective Quantities

To better visualize the evolution of the observable LTB universe, it is usually convenient to construct some “effective” parameters in a way originally used in FRW cosmology. The evolution of these parameters basically tells us what we would observe if we assumed an FRW universe, while the real universe was, say, an LTB void. The best-fit model mentioned below will be the one determined from OHD and SN Ia combined.

The first such parameter we discuss in this subsection is the effective deceleration parameter, defined to be

\[
q^\text{eff} = -1 + \frac{d\ln H}{d\ln(1 + z)}, \tag{32}
\]

which in FRW cosmology would be just the familiar deceleration parameter \( q \). In Figure 7 we show \( q^\text{eff} \) derived from the best-fit void models, compared against the deceleration parameter calculated from the best-fit flat \( \Lambda \)CDM model. The void models are seen to observationally mimic a universe that experiences an acceleration period between \( z = [0, 1] \).

The second effective quantity is the effective equation of state (EoS) of the dark energy. Before the existence of the dark
Here, for the sake of consistency, we look for the best parameter inhomogeneity. This effective parameter can be constructed mind that the dynamical properties of the dark energy people try to figure out today may really be an effect of, say, a background inhomogeneity. This effective parameter can be constructed again from $H_{\Omega}$:

$$w_{\text{eff}} = -1 + \frac{1}{3} \frac{d \ln Q}{d \ln(1+z)}. \quad (33)$$

where

$$Q = \frac{H(z)^2}{H_0(0)^2} - \Omega_0(1+z)^3. \quad (34)$$

The evolution of this effective EoS is plotted in Figure 8. Contrary to FLSC, we did not find a $-1$ crossing of the effective EoS in any of our test voids, which might be partly due to the more recent data sets of both SN Ia and OHD we have used.

### 6. FORECAST OF CONSTRAINTS FROM THE FUTURE OHD DATA SET

We saw in the last section that (1) the two-dimensional confidence regions of the parameters from OHD and SN Ia overlap (consistently), (2) the two observational probes prefer characteristically different void profiles. In this section, we explore the situation in the future with Monte Carlo simulated OHD data. The first subsection is devoted to explaining the essential ingredients of our simulation.

To facilitate the comparison between SN Ia and OHD, we introduce for each void model a characteristic parameter plane, which we refer to as CPP for brevity. The constraining power of one observational probe is quantified via the area of the confidence region on this specific parameter plane. As we will find later, this is directly related to the definition of the FoM. Take the CGBH model, for example, the two parameters $r_0$ and $dr$ contain most of the information one would like to know about the void profile, so the $r_0 - dr$ plane is chosen as its CPP. The CPPs of the Gaussian void and the Spline void are not so obvious. In fact, for the Gaussian void model, a better way might be to compare the length of the one-dimensional parameter confidence interval and, for the Spline model, the volume of the three-dimensional ($\Omega_{-2}, r_2, r_3$) parameter confidence region. Here, for the sake of consistency, we look for the best parameter pair and choose the CPP for the Gaussian and the Spline void models to be $r_0 - \Omega_0$ and $r_2 - \Omega_{r_2}$, respectively.

Based on the simulated data two questions are investigated, both of which are related to the CPP chosen above. In the following we describe them in turn. First, it is possible that when more OHD data points are obtained, they become more consistent with the SN Ia data. On the CPP, this would mean that the $1\sigma$ and $2\sigma$ confidence regions obtained from OHD become more overlapped with those from SN Ia. One would like to compare in this case the constraining power of the two probes. The first question (Question-I) is therefore raised, that when does the area of the OHD’s confidence regions on the CPP become smaller than that of SN Ia. The standard way to do this type of comparison is via the FoM. We calculate FoMs for our simulated data sets in the second subsection.

Considering the different preference in the void profiles of OHD and SN Ia, the second question (Question-II) one would like to know through the OHD data simulation is how many OHD data points are needed for the $1\sigma$ confidence region on the CPP totally detached from that of the SN Ia, if the future OHD data points hold the same void preference as the current ones. This will be the content of the third subsection.

### 6.1. Monte Carlo Simulation

The fiducial function of $H(z)$. This function sets the expectation value of the simulated OHD data. For the purpose of Question-II, where the future data points are assumed to have the same void preference as the current ones, a straightforward fiducial function of $H(z)$ comes from the best-fit void models obtained from the SJVKS data set of OHD.

In dealing with Question-I, on the other hand, where we assume the future OHD and SN Ia “converge” in their best-fit void profiles, and hence the model parameters, one natural choice is again to use the best-fit void models; only this time the model parameters are given by OHD and SN Ia combined. The problem is that the constraints from the current data sets are far from tight, as we have seen in Figure 2; the best-fit parameters are not Gaussian in many cases, and hence not quite well defined. The (flat $\Lambda$CDM) model, on the other hand, is consistent with and well constrained by all current observations, and is therefore generally the best knowledge we have about the observational universe; a fiducial model based on this model would seem to be a proper choice. However, flat $\Lambda$CDM, as do any other FRW models, differ characteristically from the void models. Using flat $\Lambda$CDM as the fiducial model, one risks introducing a systematic error, which would in turn cause a shift and/or a shape change in the parameter confidence region (Kim et al. 2004), and hence a bias in the FoM (see the next subsection). Noticing the merits and demerits of each of the two choices, we obtain the center values of the simulated data set from a mixture of the two (Robert & Casella 2004). More specifically, statistically half of the data points are drawn from the flat $\Lambda$CDM, with $\Omega_m = 0.27$ and $H_0 = 73$, which is consistent with the seven-year Wilkinson Microwave Anisotropy Probe (WMAP) and $H_0$ observation (Komatsu et al. 2011; Riess et al. 2011), and in generating the other half of the data, the void model with the combined best-fit parameters is used as the fiducial function.

The discussion of the following two ingredients of a Monte Carlo simulation should be the same when applied to Question-I and Question-II. Where the fiducial function is used for illustration, we use the best-fit void model (combined).

The uncertainty model. This model generalizes the statistical and systematic uncertainties about a future observation. There
is, however, no such specification for future OHD surveys, to
the authors’ knowledge. To work around this difficulty, Ma &
Zhang (2011) examined the SJVKS data set and proposed a
simple phenomenological uncertainty model. Here, in order
to gain more control on the simulations, we adopt an even
simpler strategy. We introduce a percentage, \( \delta \), to describe
the maximum uncertainty for each data point, namely, for a
given fiducial value \( H_i \), the expectation of the uncertainty \( \sigma_i \)
is taken randomly from the interval \([ -H_i \delta, H_i \delta ] \). This uncertainty
model is surely oversimplified: not all errors in the observations
are redshift dependent. Nonetheless, as shown in Figure 9, a
\( \delta = 25\% \) error line does acts as an envelope for the current
observational uncertainties, except at \( z = 0.48, 0.88 \), where the
data points have extraordinary errors (Stern et al. 2010b). Once
\( \sigma_i \) is obtained, the mid-point of the observational data will then
be generated from a Gaussian distribution \( N(H_i, \sigma_i^2) \). One such
sample is shown, against the SJVKS data set, in Figure 10.

**Figure 9.** \( H_i \delta \) line with \( \delta = 25\% \) forms an envelope of the errors of SJVKS data
set. Only two data points at \( z = 0.48, 0.88 \) are seen to be beyond the envelope.

**Figure 10.** One snapshot realization of the OHD data set, with the uncertainty
level \( \delta = 25\% \). The SJVKS data set (diamond with red error bar) are plotted for
comparison.

**Sample size and survey depth.** The survey depth and the sam-
ple size are two more important aspects of an observation that
highly influence the parameter determination. We denote them
as \( z_{\text{max}} \) and \( N_i \), respectively. Therefore, we have in total three
control parameters to describe a possible future observation (or
to generate one sample of data points): \( \{ \delta, z_{\text{max}}, N_i \} \). In practice,
\( \delta \) takes three values: 25\% (current precision), 15\% (according
to a prediction of OHD from Atacama Cosmology Telescope,
see Simon et al. 2005), 3\% (can now be achieved at \( z \approx 0.42, \)
see Crawford et al. 2010); \( z_{\text{max}} \) also takes three values: 1.5, 2,
2.5; and \( N_i \) takes four: 64, 32, 16, 8.

**6.2. Figure of Merit**

We define the FoM to be the reciprocal area enclosed by the
2\( \sigma \) marginalized likelihood contour on the CPP of each of our
void models. Higher FoM would then mean tighter, and hence
better constraint. The current values of FoM from OHD and SN
Ia for different void models are shown in Table 5: the values
in the OHD row are seen to be less than their counterparts in
the SN row. Question-I could be translated then to the “FoM
language” as: when will the future OHD data set give out an
FoM larger than that of the SN Ia Union2 data set?

To calculate the FoMs of the simulated data sets with fixed
\( z_{\text{max}} \) and \( \delta \), three steps are taken in turn: (1) generating a sample
of 64 OHD data points that are evenly spaced in the redshift
range \([0.1, z_{\text{max}}] \); (2) calculating the FoM based on this sample;
(3) Extracting subsamples with \( N_i = 32, N_i = 16, N_i = 8 \), and
calculating their FoMs in turn. Repeat these three steps for, say,
100 times, and we get for each \( N_i \) an array of 100 FoMs, the
median of which is taken as the typical FoM. Then we iterate
the whole process for different \( \{ z_{\text{max}}, \delta \} \) and finally get a three-
dimensional matrix, with the element being the typical FoM at
each \( \{ z_{\text{max}}, \delta, N_i \} \).

The result is shown in Figure 11. The size of the “+” symbol
is proportional to the value of the FoM, which, if greater than that
of the Union2 data set of SN Ia, is plotted in red, but otherwise
in blue.

Three void models are seen to share several very similar
features, both qualitatively and quantitatively. First, as one may
intuitively assume, the FoM increases with the sample size of
the Union2 data set of SN Ia for different void models are shown in Table 5: the values
in the OHD row are seen to be less than their counterparts in
the SN row. Question-I could be translated then to the “FoM
language” as: when will the future OHD data set give out an
FoM larger than that of the SN Ia Union2 data set?

The second point is not surprising either, that lowering the
uncertainty level \( \delta \) will increase the FoM. At \( \delta = 15\% \), a sample of
16 OHD data points are sufficient to surpass SN Union2 for all choices of \( z_{\text{max}} \).

The third point deserves a bit more discussion. One might
expect that a deeper survey will tighten the constraints. This
is not always true, however, judging from the value of the
FoMs. For example, when \( N_i = 16 \) and \( \delta = 25\% \), the FoM of
OHD with \( z_{\text{max}} = 1.5 \) is 0.209, 0.203, 0.234, respectively.
That is, higher \( z_{\text{max}} \) might lead to a lower FoM.

These phenomena exist in all three void models. One possible
reason for this to happen is that, for a given \( N_i \), especially when
\( N_i \) is small, a higher \( z_{\text{max}} \) leads to a looser redshift coverage,
which in turn leads to more freedom for the theoretical curve
to vary between the adjacent data points. The extra freedom in
the theoretical curve is realized, of course, through the extra
freedom of the model parameters, hence a lower FoM.
Figure 11. Figure of merit (FoM) in the control parameter space. In these right-handed coordinate systems, the x-axes are the error level $\delta$, the y-axes $z_{\text{max}}$, and the z-axes the size of the data sets $N_s$. The lengths of the plus symbols (+) are proportional to the values of the FoM. The symbol is plotted in red, and the line connected to it from below is solid if the FoM value is greater than that of the Union2 data set of SN Ia, and blue and dashed otherwise.

Figure 12. Normalized $\mu-z$ relation at different $H_0$. The distance modulus $\mu$ given by $H_0 = H$ are normalized by multiplying the corresponding luminosity distances by $H/H_{\text{BST}}$, where $H_{\text{BST}}$ is the best-fit $H_0$ value for each model, and other parameters all take the best-fit values determined from the Union2 data set. The difference between curves of $H = 45, 75, 90$ are well within the observational uncertainties.

6.3. 1\(\sigma\) Confidence Region Overlap

As discussed above, Question-II is motivated by the difference in the best-fit models obtained from the SJVKS data set of OHD, and from the Union2 data set of SN Ia. The logic is that if this difference holds for the future data, then one could finally tell the inconsistency from the model parameter confidence regions obtained from OHD and SN Ia; for instance, their 1\(\sigma\) confidence regions on the CPP will separate.

To simulate this, we generate new OHD data using the SJVKS-induced best-fit void models as the fiducial function. We “mark” at first those parameters encompassed by the SN Union2-induced 1\(\sigma\) contour on the CPP. Then for each simulated OHD data set, we determine its likelihood and count the number of the marked parameters lying inside its 1\(\sigma\) confidence region, $N_o$. For each simulation control parameter combination, we generate 100 samples, obtain 100 $N_o$s and log the median value. This median number then reflects the typical overlapping condition for a given $\{N_s, \delta, z_{\text{max}}\}$ combination, i.e., if it equals zero then very likely one could see the separation of the OHD- and SN Ia-induced 1\(\sigma\) confidence regions on the CPP.

The results from different void models are not as homogeneous as in the last subsection. The Gaussian void model is the easiest for the inconsistency to emerge: most of the control
parameter combinations, except those with \( N_s = 8, \delta = 25\% \), have a median \( N_\delta \) equal to 0. For the CGBH model, only at the largest data set size (\( N_s = 64 \)), where the highest redshift reaches \( z_{\text{max}} = 2.5 \) and the highest precision (\( \delta = 3\% \)), does the median value of \( N_\delta \) equal zero. In the case of the Spline void, \( N_\delta \) is nonzero for all the control parameter combinations, that is, from \( r_2 - \Omega_{m2} \) plane at least SN Ia and the OHD are always consistent. Therefore, the prediction about when the OHD observation is powerful enough that, when combined with SN Union2 data set, it can help rule out the void models, depends on the specific void model. This is our answer to Question-II.

7. CONCLUSIONS

We have discussed in this work the use of OHD, both in the present and in the future, as a probe to constrain the LTB models.

We first showed, that of the two expansion rates involved in LTB models, the one in the radial direction corresponds to the observables in OHD, which are obtained from the passively evolving galaxies. The latest OHD compilation, also the one discussed in this work, is the SJVKS data set updated and compiled in Stern et al. (2010a). A subtlety is then discussed on using OHD in the LTB scenario: on the one hand, a synchronized oldest-galaxies forming time throughout the universe is the basic assumption behind the OHD observation, but on the other hand, LTB models have large-scale background inhomogeneities, and there is no good reason to believe that the first galaxies formed simultaneously all over the universe.

The problem is severe because if the data set itself is problematic, then surely it cannot be used to constrain any model. We have examined the method in obtaining OHD and found that only the homogeneity inside the redshift bin \((<0.15)\) where each data point is obtained is demanded for the data to be valid. This is not an unreasonable assumption for the LTB models. In the standard Onion approximation, the LTB models are approximated as being composed of different spherical layers, inside which the matter is treated as homogeneously distributed and the dynamics are FRW-like. However, whether a redshift bin can be viewed as one such layer should be checked given the model parameters. We argue, therefore, that a complete parameter constraining analysis should have a consistency check procedure after the standard \( \chi^2 \) analysis, and we illustrated this process on our test models with the SJVKS data at the end of Section 4. The result is inconclusive due to the large uncertainties in the absolute ages of the oldest galaxies.

The constraints obtained from the likelihood analysis are indicative. SN Ia performs generally better than OHD judging from the likelihood confidence regions, but the two observational probes both favor the voids with the center largely depleted of matter. Besides, and more remarkable still, although the 1 or confidence regions from OHD overlap with that from the SN Ia data in all of the test models, the former favors the broad and flat voids, while the latter prefers voids that are smaller and more steep around the center.

In Section 5 we compared the best-fit void models with the best-fit flat ΛCDM, first through the corrected AICc analysis. The latter, according to the AICc value, is better than any of our void models. Indeed, especially for the complicated models, the small size of the SJVKS data set of OHD greatly degrades their performance in the AICc analysis.

The dynamical properties are then compared in terms of the effective deceleration parameter (\( q^{\text{eff}} \)) and the effective EoS of dark energy (\( w^{\text{eff}} \)). The \( q^{\text{eff}} \) is seen to enter the negative region recently (\( z < 1 \)) for all the void models, mimicking the recent acceleration of the universe. The \( w^{\text{eff}} \) of the void models are negative today, mimicking the dark energy’s repulsive power.

This general consistency and the characteristic difference in the best-fit voids of SN Ia and OHD, obtained in Section 4, are taken further by using the Monte Carlo simulated data in Section 6. More specifically, we use the simulated data to predict when OHD become more constraining than the Union2 data set of the SN Ia (Question-I), and when the two probes become inconsistent with each other (Question-II), respectively.

We made use of the FoM analysis for Question-I, and the answer turned out to be surprisingly homogeneous for the three models: \( N_s = 32 \) is needed for the FoM of OHD to be greater than that of the Union2 data set when \( \delta = 25\% \); \( N_s = 16 \) when \( \delta = 15\% \); and \( N_s = 8 \) when \( \delta = 3\% \); increasing \( z_{\text{max}} \) does not necessarily increase the constraining power.

The answer to Question-II is not so simple and is only semi-quantitative: to observe an inconsistency for the Gaussian void model, a data set with slightly more data points than SJVKS (\( N_s = 16 \)), and at the current error level (\( \delta = 25\% \)), will do the job, but only when \( N_s, \delta, \) and \( z_{\text{max}} \) are all at their most demanding level (64, 3%, and 2.5, respectively) does the inconsistency show up; in the model with the most parameters, the Spline void model, the simulated OHD and SN Ia are consistent for all the possible \( \{ N_s, \delta, z_{\text{max}} \} \) combinations.

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APPENDIX A

\( H_0 \) AS A NORMALIZATION FACTOR

The observed distance modulus data is given by (Guy et al. 2007; Amanullah et al. 2010)

\[
\mu^\text{obs} = m_B^\text{obs} - M + \alpha_s \cdot x_1 - \beta \cdot c, \quad (A1)
\]

where \( m_B, x_1, c \) are derived from the fit to the supernovae light curves and therefore are known before fitting the cosmology, \( M, \alpha_s \), and \( \beta \) are fitted together with the cosmological parameters to minimize the residuals in the Hubble diagram.

The theoretical distance modulus \( \mu \) is derived from the luminosity distance,

\[
\mu^\text{th} = 5 \log \left( \frac{d_L}{\text{Mpc}} \right) + 25. \quad (A2)
\]

In FRW cosmology, \( d_L \) is proportional to the reciprocal of the Hubble parameter,

\[
d_L^\text{FRW} = (H_0^\text{FRW})^{-1} D_L(\theta), \quad (A3)
\]

where \( D_L \) is a function of the cosmological components, \( \theta \). One can see how \( H_0^\text{FRW} \) totally degenerates in this case with the absolute magnitudes of the supernovae in Equation (A1) by
Figure 13. Constraints on Gaussian void model parameters from OHD and SN Ia. The red and brown filled regions correspond, respectively, to 1σ and 2σ confidence regions obtained from the SJVKS data set of OHD. Overplotted are the 1σ and 2σ confidence regions obtained from the Union2 data set of SN Ia, in solid and dotted lines, respectively.

In LTB models, the expansion rate goes into the distance in a more complicated way. When keeping the \( \Omega(r) \) profile unchanged and varying \( H_0 \), the pattern of distance-modulus–redshifts curves also changes, as is argued and shown in Figure 4 of FLSC. However, we found that in the redshifts range we are concerned with, \( z \leq 2.5 \), treating \( H_0 \) as a normalization factor in all our testing models leads to an error that is too small to be of any importance.

More specifically, we calculate the luminosity distance for each model at their best-fit parameters, then redo the calculation with the value of \( H_0 \) changed to be 45 (and then 90) and the other parameters unchanged. The “normalized” luminosity distances

\[
\chi^2(M, \alpha, \beta, H_0^{FRW}, \theta) = \sum_i \frac{\mu^{FRW}(M, \alpha, \beta, H_0^{FRW}, \theta, z_i) - \mu^{obs}}{\sigma_i^2}
\]

\[
= \sum_i \frac{m_i^2 - 5 \log_{10} D_L(\theta, z_i) - 25 + (M + \alpha \cdot x_1 - \beta \cdot c + 5 \log_{10} H_0^{FRW})}{\sigma_i^2}
\]

This makes it impossible to obtain constraints of \( H_0^{FRW} \) from, say, the Union2 data set in the FRW framework.
are calculated to be
\begin{align}
d_L^{N45} &= \frac{d_L^{05}}{R_{G}^{05}}, \\
d_L^{N90} &= \frac{d_L^{00}}{R_{G}^{00}},
\end{align}
respectively, where the superscript “BST” denotes the best-fit value. The normalized distance modulus $\mu^N$ is then obtained according to Equation (A2). The two specific numbers, 45 and 90, are so chosen that the values of $H_0$ scanned in our calculation all lie between them; therefore, this choice should serve as an upper bound of the possible error introduced by the approximation. The relative difference in the normalized distance modulus varying $H_0$ from 45 to 90 is
\begin{equation}
\delta_{\mu} = \frac{\mu^{N90} - \mu^{N45}}{\mu^{N45}}.
\end{equation}
The maximum value of $\delta_{\mu}$ at $z \leq 2.5$ for the CGBH model is 0.54%, Gaussian void 0.26%, and Spline void 0.05%, respectively, i.e., all less than 1% and cannot be differentiated by the current observation (see Figure 12 for an illustration). Therefore, we conclude that, at least for the models discussed and the redshifts concerned in the current work, $H_0$ can still be regarded as a normalization factor.

APPENDIX B
CONSTRAINTS ON THE GAUSSIAN AND THE SPLINE VOID MODEL

The constraints on the Gaussian and the Spline void models are shown in Figures 13 and 14, respectively. The scanning ranges of the parameters are the same, as can be told from the axis in the plots. From the constraints, one can tell why the answers to Question-II in Section 6 are so different for the Gaussian void and the Spline void models: in the Gaussian void model the $1\sigma$ regions of $r_0 - \Omega_0$ from OHD and SN Ia are only slightly overlapped, so a few more data points of OHD will be able to separate them; the confidence regions of $r_2 - \Omega_2$ of the Spline void model, on the other hand, contain a much larger mutual part, therefore many more data points and/or much higher precision are needed to separate them.

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