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Development and evaluation of the psychometric properties of a brief parenting scale (PS-7) for the parents of adolescents

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Abstract

This study aimed to develop a seven-item brief parenting scale (PS-7) based on the original parenting scale (PS) and various other shortened versions and with a better factor structure for the parents of adolescents. The scale was tested with a sample of 3,777 parents (2,205 mothers and 1,572 fathers). Confirmatory factor analysis was used to evaluate the dimensionality of the different versions of the PS. Only the PS-7 with a two-factor structure, i.e., laxness (three items) and overreactivity (four items), showed a good model fit based on a representative sample of parents of junior secondary school students. Overall, the results suggest that PS-7 is comparable to the original PS and possesses good psychometric properties in terms of internal consistency, factorial validity, construct validity, criterion validity and discriminant validity. The abbreviated parenting scale also provides a reliable and cost-effective method for assessing parental practices for treatment and assessing treatment outcomes.

Background

Parenting has been shown to have an important influence on the mental well-being of children and adolescents and the prevalence of behavioural problems [1–4]. The literature has also suggested that parenting influences children’s school performance and has linked parenting styles to controversies arising over cultural differences. Asian, and especially Chinese, parenting styles have been categorised as controlling and authoritarian, and are popularly known as ‘tiger’ parenting [5]. However, a longitudinal study of 444 Chinese American families that examined the effects of parenting styles on adolescent adjustment suggested that tiger parenting was not the most typical profile [6]. Going beyond the common perception of Asian parenting as controlling and authoritarian [7–10], recent studies have suggested that close parental control and an authoritative parenting style are fused with ideas of training and presence that help to explain school achievement [11, 12]. Parenting style has now been accepted as a cross-cultural concept that enhances understanding of child behaviour across ages and ethnicities [7, 13, 14]. Hence, developing a convenient cross-cultural measure of parenting styles is highly warranted for epistemological research.
To operationalise parenting style, Arnold, O’Leary [15] developed a parenting scale (PS) based on a sample of 168 mothers of children ranging from 18 to 48 months old (98 boys and 70 girls). Subsequent studies suggested that the scale was applicable not only to mothers of toddlers, but also to parents of both genders with children and young adolescents attending primary and secondary schools [16–21]. The PS has since been widely accepted internationally as a measure of parenting behaviour [16]. The original and adapted versions of the scale have been translated into numerous languages, including Chinese [17], Dutch [18], French [22], German [16, 19], Japanese [23, 24], Persian [19], Spanish [25], Swedish [20] and Vietnamese [21]. The scale has also been used to examine the behaviour of parents in different contexts, such as community-based paediatric practices for routine care in America [26], Australian mothers with preschool-aged children [27], parents of school-aged children with ADHD [28] and clinical populations [29].

Nevertheless, several factors may limit the application of the full version of the PS. The scale originally comprised 30 items with a three-factor structure, comprising laxness (11 items), overreactivity (10 items) and verbosity (7 items; with two multi-factor items, 7 and 9). The scale developers reported that four items (1, 5, 13 and 27) with low factor loading values (below 0.35) were categorised as not loading on a specific factor and were excluded from the scale. Hence, the 26 item PS with a three-factor structure is commonly used. The original scale was derived based on exploratory factor analysis (EFA), and ambiguous results were obtained for the dimensionality and number of items per factor. In particular, verbosity was found to have a complicated factor structure and coefficients with a questionable alpha value of 0.63 [15]. Moreover, the numerous studies conducted during the early development and application of the scale mainly focused on relatively small samples of mothers with infants and English-speaking populations [18, 30].

To address these issues, many early studies attempted to provide a shortened version of the PS [31]. However, these studies used limited validation tools to evaluate the latent structure of the scale, such as EFA to uncover the underlying structure or confirmatory factor analysis (CFA) to verify the factor structure [20]. The following five brief versions are the most significant examples and have been widely used in the field. Salari, Terreros [20] proposed a 21-item scale (PS-21) in which all of the verbosity items were removed and the original two sub-scales, laxness (11 items) and overreactivity (10 items), were included after evaluating the psychometric properties of the scale. However, the CFA failed to fulfil the goodness-of-fit indices, i.e., chi-square divided by less than or equal to three degrees of freedom or a comparative fit index (CFI) higher than 0.950 [32–34]. One of the original PS scale developers, Susan O’Leary, and her colleague proposed a 13-item shortened version of the scale (PS-13) with a three-factor structure comprising laxness (five items), overreactivity (five items) and hostility (three items) [30]. However, their newly proposed factor, hostility, had a problematic Cronbach’s alpha value of 0.52. Irvine, Biglan [35] developed a version of the PS for adolescents (PS-12), based on a sample of 298 parents (94.5% mothers) of school students who identified as being at risk for problem behaviour. The 12 items were derived from the original PS sub-scales for laxness (six items) and overreactivity (six items). Intriguingly, without the support of EFA or CFA, the authors further suggested adding an additional single monitoring item, i.e., item 13, which had been removed from the original PS scale due to low factor loading. Another shortened version was based on the findings of two studies on 187 and 216 American mothers, which suggested using a 10-item PS (PS-10) with a two-factor structure comprising laxness (5 items) and over-reactivity (5 items) [31]. Nevertheless, because the studies focused solely on mothers, the results may have limited applicability to fathers. Finally, the latest attempt was the eight-item parenting scale short form (PS-8), which comprised two sub-scales, laxness (four items) and overreactivity (four items), derived from a sample of 539 German parents (312 mothers and
This study has two main aims. First, to evaluate the factor structure of the full PS and variants of the shortened versions using CFA and a larger sample comprising the parents (both fathers and mothers) of adolescents. Second, to propose a seven-item brief parenting scale (PS-7) that has a better factor structure and better psychometric properties than the existing versions.

Methods

This study was approved by the ethical committee of the City University of Hong Kong. Its procedure was in compliance with the Declaration of Helsinki guidelines. All of the participants gave informed consent prior to the study.

Participants and measures

In January 2018, 4,007 respondents from 10 secondary schools located in different districts of Hong Kong were recruited to participate in this cross-sectional study. Respondents who were either the father or mother of an adolescent were included in the analysis (N = 3,777). The valid sample consisted of 2,205 mothers and 1,572 fathers (average age 44.83 years; SD = 6.95) of junior secondary school students (i.e., Forms 1 to 3) aged between 12 to 14. The demographic information of the participants is summarised in Table 1. The unique historical context of Hong Kong, with its mix of Eastern and Western cultures, provides an ideal research setting for investigating parenting styles because it may generate results that are relevant not only to Chinese society, but also to other Anglo-Saxon societies.

The full PS consists of 26 items with a three-factor structure comprising 11 items related to laxness (7, 8, 12, 15, 16, 19, 20, 21, 24, 26 and 30), 10 items for overreactivity (3, 6, 9, 10, 14, 17, 18, 22, 25 and 28) and 7 items for verbosity (2, 4, 7, 9, 11, 23 and 29). There are two multi-factor items: item 7, which is related to both laxness and verbosity, and item 9, which is associated with overreactivity and verbosity. The parents rated the items on a 7-point Likert scale to indicate their tendency to use specific strategies to discipline their children. The scale items were translated into Chinese using the back-translation procedure by two bilingual translators who were familiar with both Chinese and English and were fully aware of the issues and techniques relating to cross-cultural research.

Item selection process

The process is based on the criteria, the latest practice and recommendations used in the existing PS studies and other scale development and validation literature. The selected items have gone through the following two-step procedure. Step one, selecting the items: i) using inductive approach to analyze the correlation matrix of all the items and keeping the items with 0.250 or above. We also cross-checking the Cronbach’s alpha, if deleted and McDonald’s omega values to ensure that the shortened version is above the acceptable range > 0.70; ii) using scree test in factor analysis to identify the factor structure with eigenvalues higher than 1.0. We also select the items with highest factor loadings, i.e. > 0.50 and avoid items involve correlating the error terms based on the modification indices. When selecting the items, we try to retain the sufficient items (at least three) in each factor to ensure that the validity standard of the shortened version is equivalent to the full version; iii) to verify the abbreviated version with the confirmatory factor analysis to ensure that the scale with good construct validity, i.e. fulfil all the stringent requirements for good model fit. Step two,
ensuring that the compatibility between the full scale and shortened version: iv) we adopted the following practice of Kliem, Lohmann [16], ‘short form should also correlate strongly with the original PS on the total score level as well as on the subscale (overreactivity and laxness) level’ (p. 34). As such, there should be significant strong positive correlation (> 0.80) between the full and short scales, including their sub-scales; and v) lastly, the abbreviated version should possessing good criterion validity as reported in the existing PS literature.

### Procedure

The sample (N = 3,777) was randomly stratified into three datasets (samples 1, 2 and 3). Each sub-sample consisted of 1,259 cases that reflected the original sex ratio of the participants, i.e., mothers 58.4% and fathers 41.6%, to avoid the problem of overfitting when using EFA and CFA to evaluate the factorial and construct validity of the scale [51, 52].

Various psychometric testing tools and validated instruments were used to examine the newly proposed PS-7. EFA was used to evaluate the factorial validity and the principal axis method with oblique rotation was used to evaluate the factor structure of the scale [18, 34, 53]. In addition, the Kaiser-Meyer-Olkin (KMO) test and Bartlett’s test of sphericity were used to evaluate the model sufficiency. The KMO estimates were over 0.70 and the Bartlett’s test was significant (p < 0.01), thus indicating that the scale had a satisfactory factor structure [54]. According to Hair [34], an item with a factor loading over 0.50 is regarded as having practical

### Table 1. Participant demographic characteristics.

| Variable                       | Respondents |
|--------------------------------|-------------|
| Filler’s age mean (SD)         | 44.83 (6.95)|
| Partner’s age mean (SD)        | 45.35 (7.03)|
| Relationship with the target child n (%) |             |
| Mother                         | 2,205 (55%) |
| Father                         | 1,572 (39.2%) |
| Others                         | 160 (3.9%)  |
| Missing                        | 70 (1.8%)   |
| Children school year           |             |
| Form 1 n (%)                   | 1,110 (27.7%) |
| Form 2 n (%)                   | 1,150 (38.7%) |
| Form 3 n (%)                   | 1,347 (33.6%) |
| Number of children (SD)        | 2.26 (0.98) |
| Education level n (%)          |             |
| No formal education            | 973 (24.3%) |
| Primary education              | 1,520 (37.9%) |
| Secondary education            | 1,096 (29.0%) |
| Diploma or college             | 58 (2.0%)   |
| Tertiary education             | 116 (2.9%)  |
| Missing                        | 155 (3.9%)  |
| Martial status n (%)           |             |
| Single                         | 46 (1.1%)   |
| Married                        | 3,295 (82.2%) |
| Divorce/separated              | 288 (7.2%)  |
| Cohabit                        | 106 (2.6%)  |
| Widowed                        | 107 (2.7%)  |
| Missing                        | 165 (4.1%)  |

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significance in studies with over 350 respondents. The internal consistency of the scale was assessed by Cronbach’s alpha [55], McDonald’s omega [56–58] and the corrected item-total correlation between the seven items [34, 59].

CFA was used to replicate and evaluate the construct validity of the scales [42, 60, 61]. Diagonally weighted least squares (DWLS) was used as the CFA estimator to examine the factor structure of the PS for two reasons. First, the literature suggests that the PS has high item-level skewness and kurtosis [30]. Second, because scales with latent constructs estimated by Likert scale items consist of ordinal data, DWLS is regarded as the least biased and most optimal fit [62–66]. The model fit and cut-off criteria were evaluated on the basis of the values suggested in the structural equation modelling (SEM) literature. Specifically, over 0.950 for both CFI and the Tucker-Lewis fit index (TLI), below 0.08 for the standardised root mean square residual (SRMR) and below 0.06 for the root mean square error of approximation (RMSEA) are considered to indicate a good fit [32, 34, 67, 68]. In addition, model acceptability was indicated by $\chi^2 / df \leq 3$ due to the large sample size [33, 69].

The criterion validity was evaluated using other validation constructs and measurements reported in the literature on parenting. The PS has been reported to be significantly positively related to aggressive and delinquent behaviours [30, 31, 35], authoritative parenting [31], ADHD and cognitive and hyperactivity symptoms [30]. Hence, the following well-established scales were used to evaluate the criterion validity of PS-7. The reactive–proactive aggression questionnaire (RPQ) comprises 23 items to measure reactive (11 items) and proactive (12 items) forms of aggression on a Likert-type scale ranging from 0 = never to 2 = usually [70–72]. The child behaviour checklist (CBC) consists of 33 items identifying aggressive (20 items) and delinquent (13 items) behaviours on a 3-point Likert-type scale ranging from 0 = unsuitable to 2 = very suitable [73–76]. Conners’ parent rating scale (CPRS) comprises 28 items with a 4-point Likert-type scale (0 = never; 4 = a lot) for parents to rate their child in four dimensions, namely ADHD, oppositional, cognitive problems and hyperactivity [77, 78]. The parenting styles and dimensions questionnaire (PSDQ) is evaluated on a 5-point Likert-type scale (1 = never; 5 = always), with a particular focus on the three dimensions of physical coercion (five items), punitive (three items) and verbal hostility (three items) [13, 79].

In addition, Reitman, Currier [31] found that the original PS was not correlated with the educational level of the parent, and this study attempted to replicate this finding to demonstrate the discriminant validity of PS-7 [80]. The above analyses were all implemented with IBM SPSS 25.0 and the lavaan package version 0.6–3 [81] in R computing environment 3.6.0.

Results
Development of the seven-item brief parenting scale using EFA

The seven-item parenting scale was inspired by PS-12 [35], PS-10 [31] and PS-8 [16]. The selection of items for the brief version adhered to the existing practices recommended in the scale development and validation literature, with a particular focus on the cultural context and the results of inter-item correlations, corrected item-total correlations, Cronbach’s alpha, McDonald’s omega and EFA [42, 51]. The detail item selection procedure and criteria have been stated in the methods section. According to the results, the newly proposed PS-7 has a two-factor structure comprising laxness (items 16, 20 and 30) and overreactivity (items 6, 10, 14 and 17) (see the S1 Appendix). The KMO test (0.823) and Bartlett’s test of sphericity ($\chi^2 = 2452.585, p < .001$) factor analysis results from sample 1 (n = 1,259) indicate that PS-7 has an appropriate scale construction. The EFA results using the oblique rotation method (Table 2)
suggest that the two factors extracted with eigenvalues greater than 1.0 (3.123 for items related to laxness and 1.172 for items related to overreactivity) from PS-7 account for 62.195% of the total variance. The items related to laxness explain 45.708% of the variance, with factor loadings ranging from 0.733 to 0.857. The overreactivity items, which have factor loadings ranging from 0.747 to 0.801, explain 16.487% of the variance. The EFA results replicate the latent structure of the two factors, namely laxness and overreactivity, as suggested in the PS literature [15, 18, 35].

**Internal consistency**

Table 3 presents the descriptive statistics and item correlations for all seven items of PS-7 from sample 1. The corrected item-to-total correlations for PS-7 range from 0.470 to 0.599, which is similar to the range of 0.42 to 0.65 reported by Kliem, Lohmann [16]. Cronbach’s alpha for the seven-item PS (0.799) is comparable to that reported by Kliem, Lohmann [16] (0.75) and to the values reported in other related studies. McDonald’s omega (0.83) also suggests that PS-7 has good internal consistency.
Factor structure and comparison with other PS constructs

The factor analysis results for sample 2 (n = 1,259) replicate the findings of sample 1. The KMO test and Bartlett’s test of sphericity values are 0.827 and chi-square = 2229.075 (p < .001), respectively. The newly proposed PS-7 records 60.716% of the total variance explained by the EFA with oblique rotation. The overreactivity items (6, 10, 14 and 17) have factor loadings ranging from 0.695 to 0.905 and explain 44.614% of the variance. The laxness items (16, 20 and 30) with λ = 0.750 to 0.827 explain 16.102% of the variance. The coefficient alpha of PS-7 (0.790) in sample 2 is also above the acceptable level.

Table 4 shows the CFA results (sample 2; n = 1,259) for the original PS [15] and various shortened versions suggested in the literature [16, 20, 30, 31, 35]. All of the models evaluated in this study are without correlating measurement errors. The CFA results suggest that none of the above scales meet the minimum criteria for adequate or good model fit. The results for the original PS scale are \( \chi^2 (4979.560) / 294 = 16.94 \), SRMR = 0.086 and RMSEA = 0.113. The other four shortened versions [20, 30, 31, 35] also fail to obtain a satisfactory model fit, with either the \( \chi^2/df \) or RMSEA values being too low. The CFA results for the latest PS-8 version proposed by Kliem, Lohmann [16] satisfies all of the cut-off values for good fit other than \( \chi^2/df > 3 \).

The CFA results indicate that PS-7 has good model fit, with \( \chi^2 (21.809) / 13 = 1.68, p = 0.058 \), SRMR = 0.020, CFI = 0.999, TLI = 0.998 and RMSEA = 0.023. The standardised factor loadings for the CFA results are high, ranging from 0.64 to 0.77. Overall, the results indicate that PS-7 generally has good fit for a two underlying factor structure without any post hoc modifications.

Construct validity

This section further evaluates the psychometric properties of PS-7 with reference to the construct validity based on the data from samples 2 (n = 1,259) and 3 (n = 1,259). The CFA results in Table 5 (see Fig 1 for estimated model) suggest that all of the models fulfil the criteria for good model fit. In particular, the results for sample 3 (\( \alpha = 0.79; \omega = 0.84 \)) are \( \chi^2 (16.729) / 13 = 1.29, p = 0.212 \), SRMR = 0.017, CFI = 0.999, TLI = 0.999 and RMSEA = 0.015. The results support the two-factor structure of PS-7, i.e., laxness (items 16, 20 and 30) and overreactivity (items 6, 10, 14 and 17).

Criterion validity and discriminant validity

Table 6 shows that PS-7 is strongly correlated with the original PS-26 in terms of the total score and the subscales, namely overreactivity and laxness, for the entire sample (N = 3,777). PS-7 is very significantly positively correlated (\( r = 0.916, r_s = 0.915, p < 0.001 \)) with PS-26 and

Table 4. Confirmatory factor analysis of the parenting scale.

| Model [No. of factor/item] | \( \chi^2 \) | df | \( \chi^2/df \) | RMSEA [90% CI] | CFI | TLI | SRMR |
|---------------------------|-------------|----|----------------|----------------|-----|-----|------|
| Aronld et al., (1993) [3/26] | 4979.560 | 294 | 16.94 | 0.113 [0.110–0.116] | 0.958 | 0.953 | 0.086 |
| Salari et al., (2012) [2/21] | 2277.121 | 188 | 12.11 | 0.094 [0.091–0.098] | 0.975 | 0.972 | 0.069 |
| Rhoades & O’Leary, (2007) [3/13] | 554.963 | 62 | 8.95 | 0.080 [0.074–0.086] | 0.987 | 0.983 | 0.053 |
| Irvine et al., (1999) [2/12] | 254.697 | 53 | 4.81 | 0.055 [0.048–0.062] | 0.994 | 0.992 | 0.038 |
| Reitman et al., (2001) [2/10] | 265.459 | 34 | 7.81 | 0.074 [0.066–0.082] | 0.988 | 0.984 | 0.047 |
| Kliem et al., (2019) [2/8] | 73.870 | 19 | 3.89 | 0.048 [0.037–0.060] | 0.995 | 0.993 | 0.031 |
| PS-7 [2/7] | 21.809 | 13 | 1.68 | 0.023 [0.000–0.040] | 0.999 | 0.998 | 0.020 |

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its sub-scales, i.e., laxness ($r = 0.830, r_s = 0.817, p < 0.001$) and overreactivity ($r = 0.850, r_s = 0.852, p < 0.001$).

The results presented in Table 7 replicate the relationship between PS-7 and the other construct-related scales suggested in the literature [30, 31, 35]. The CBC aggressive and delinquent dimensions are significantly moderately correlated with PS-7 and the laxness and overreactivity subscales. The parents’ reports on reactive and proactive aggression are also positively correlated with PS-7, with $r = 0.318 (p < 0.001)$ and $r = 0.249 (p < 0.001)$, respectively. PS-7 is

### Table 5. Factor loadings and fit indices in confirmatory factor analysis for the PS-7, by sample (see Fig 1 for estimated model).

| Factor/question number | Sample 2 | Sample 3 | Combo |
|-------------------------|----------|----------|--------|
|                         | Mother   | Father   | All    |
| Laxness (LAX)           |          |          |        |
| 16 $\lambda_1$         | 0.653    | 0.607    | 0.640  |
| 20 $\lambda_2$         | 0.726    | 0.727    | 0.737  |
| 30 $\lambda_3$         | 0.821    | 0.760    | 0.768  |
| Overreactivity (OVE)    |          |          |        |
| 6 $\lambda_4$          | 0.706    | 0.701    | 0.686  |
| 10 $\lambda_5$         | 0.679    | 0.664    | 0.695  |
| 14 $\lambda_6$         | 0.718    | 0.764    | 0.701  |
| 17 $\lambda_7$         | 0.754    | 0.767    | 0.775  |
| Latent factor covariance|          |          |        |
| Laxness ~ Overreactivity| $\phi_{lo}$ | 0.614    | 0.623  |

**Model fit**

| N   | Mother | Father | All |
|-----|--------|--------|-----|
| RMSEA | 0.000  | 0.050  | 0.030 |
| RMSEA 90% CI | 0.000–0.030 | 0.000–0.053 | 0.000–0.037 |
| SRMR  | 0.018  | 0.025  | 0.020 |
| $\chi^2$ (df = 13) | 10.331 | 17.219 | 21.809 |
| $\chi^2$/df | 0.79   | 1.32   | 1.68  |
| CFI   | 0.999  | 0.999  | 0.999 |
| TLI   | 0.998  | 0.998  | 0.998 |

Combo = sample 2 plus sample 3 (n = 2,518)

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**Fig 1. Estimated model of the 7-item parenting scale.**

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also significantly correlated with authoritarian parenting styles, such as physical coercion \( (r = 0.383, p < 0.001) \), punitive behaviour \( (r = 0.399, p < 0.001) \) and verbal hostility \( (r = 0.495, p < 0.001) \). The parents reported that their children manifested emotional and behavioural symptoms, including ADHD \( (r = 0.354, p < 0.001) \), oppositional behaviour \( (r = 0.388, p < 0.001) \), cognitive problems \( (r = 0.315, p < 0.001) \) and hyperactivity \( (r = 0.355, p < 0.001) \). This also correlates with the shortened version of the PS. The results also replicate the finding that PS-7 is not significantly related to the educational level of the parent [31], with the results showing that \( r = -0.009 \) \( (p = 0.581) \), PS-7: laxness \( r = -0.016 \) \( (p = 0.333) \) and PS-7: overreactivity \( r = 0.001 \) \( (p = 0.963) \). Thus, PS-7 generally has good criterion and divergent validity.

Discussion

The main contribution of this study is to introduce PS-7, a shortened version of the original PS. PS-7 and its sub-scales have very strong significantly positive relationships with the

Table 6. Correlations for the PS-7 and PS-26 sub-scales (N = 3,777).

| Scale                  | (1)   | (2)   | (3)   | (4)   | (5)   | (6)   |
|------------------------|-------|-------|-------|-------|-------|-------|
| 1. PS-7: total score   | 1.000 | 0.814*** | 0.885*** | 0.916*** | 0.830*** | 0.850*** |
| 2. PS-7: laxness       | 0.804*** | 1.000 | 0.450*** | 0.714*** | 0.878**  | 0.462*** |
| 3. PS-7: overreactivity| 0.885*** | 0.453*** | 1.000 | 0.836*** | 0.573*** | 0.937*** |
| 4. PS-26: total score  | 0.915*** | 0.700*** | 0.844*** | 1.000 | 0.863*** | 0.878*** |
| 5. PS-26: laxness      | 0.817*** | 0.877*** | 0.566*** | 0.844*** | 1.000 | 0.589*** |
| 6. PS-26: overreactivity| 0.852*** | 0.467*** | 0.940*** | 0.887*** | 0.584*** | 1.000 |

* \( p < .05 \).
** \( p < .01 \).
*** \( p < .001 \).

Lower triangle for Spearman correlations; upper triangle for Pearson correlations

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Table 7. Correlations between the PS-7 in relation to other construct-related scales (N = 3,777).

| Scale                          | PS-7   | PS-7: Laxness | PS-7: Overreactivity |
|-------------------------------|--------|---------------|----------------------|
| Criterion validity            |        |               |                      |
| CBC: Aggressive               | 0.347***| 0.220***      | 0.358***             |
| CBC: Delinquent               | 0.304***| 0.183***      | 0.320***             |
| RPQ-parent-report: reactive   | 0.318***| 0.195***      | 0.335***             |
| RPQ-parent-report: proactive  | 0.249***| 0.154***      | 0.259***             |
| PSDQ: physical coercion       | 0.383***| 0.163***      | 0.456***             |
| PSDQ: punitive                | 0.399***| 0.204***      | 0.450***             |
| PSDQ: verbal hostility        | 0.495***| 0.223***      | 0.581***             |
| Parent rating: ADHD           | 0.354***| 0.228***      | 0.360***             |
| Parent rating: Oppositional   | 0.388***| 0.243***      | 0.402***             |
| Parent rating: Cognitive problem| 0.315** | 0.196**       | 0.327**              |
| Parent rating: Hyperactivity  | 0.355***| 0.219***      | 0.368***             |
| Divergent validity            |        |               |                      |
| Parent’s educational level    | -0.009 | -0.016        | -0.001               |

* \( p < .05 \).
** \( p < .01 \).
*** \( p < .001 \).

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original scale and its sub-scales, which suggests that PS-7 is comparable to the original PS. The proposed scale retains the original two-factor structure, i.e., laxness and overreactivity, as suggested in the PS literature [16, 20, 31, 35]. The shortened scale also demonstrates good psychometric properties in terms of internal consistency and factorial, criterion and discriminant validity. Thus, the proposed PS-7 provides a handy instrument for researchers and practitioners wishing to evaluate parenting practices for fathers and mothers of young adolescents.

PS-7 is preferable to the existing versions for the following reasons. First, the adapted scale has no complicated items and possesses better factorial validity, with the CFA results suggesting an excellent model fit. Second, in some studies, only EFA and Cronbach’s alpha were used to evaluate the metrics of the scales [15, 20]. In this study, the proposed PS-7 was subjected to a series of rigorous tests and comprehensive psychometric tools were used to develop and validate the scale. The results showed that PS-7 has a better factor structure than and is comparable to the original PS. Finally, PS-7 does not rely on correlating the error terms to fulfill the stringent requirements of the goodness-of-fit in CFA. Nonetheless, the proponents of the existing PS versions largely relied on modification indices to improve the model fit [16, 18, 30, 35]. According to Hermida [82], it is inappropriate to allow correlated errors in SEM without strong theoretical justification. Hence, PS-7 is more favourable than the existing PS versions.

Some PS items were not included in PS-7 mainly due to concerns about cultural sensitivity and the contextual rules and regulations. The notion of paternalism is deeply embedded in Asian societies [83, 84]. Therefore, item 12 (When I want my child to stop doing something, I coax or beg my child to stop) and item 21 (If saying “No” doesn’t work, I offer my child something nice so he/she will behave) are less likely to be relevant in an Asian context when parents interact with their children. Similarly, the scenario in item 22 (When my child misbehaves, I get so frustrated or angry that my child can see I’m upset) is unlikely to arise in Chinese society because the notion of face prevents parents from showing any signs of weakness in front of their children [85]. In many societies, including Hong Kong, laws and regulations forbid parents imposing physical punishment and leaving their children unattended at home [86]. Therefore, item 15 (When we’re not at home, I let my child get away with a lot more) and item 18 (When my child misbehaves, I spank, slap, grab, or hit my child) may not be applicable in those societies. Future studies should consider the significance of such cultural differences.

A potential limitation of this study is that only a limited number of construal-related scales were used to evaluate the criterion validity of PS-7. In the PS literature, the scales are normally cross-checked with measures such as depression, anxiety, self-esteem, confidence, parent-child relationship, impulsivity and social support [16, 30, 31, 35]. Due to the availability of Chinese validated scales and to avoid a lengthy questionnaire, this study used other well-developed scales related to children’s aggressive and delinquent behaviour, authoritative parenting, ADHD and oppositional, cognitive and hyperactivity symptoms, which have been extensively discussed and used in the PS literature. The sample used in this study may also limit the generalisability of the findings given that the respondents were recruited from junior secondary schools in Hong Kong and the lack of any evaluation of test-retest reliability. However, these limitations may have been compensated by the large sample size and inclusion of father and mother respondents. Further research is needed to replicate our findings or apply PS-7 in other contexts, preferably with cross-cultural longitudinal research designs in different societies, and ideally involving fathers and mothers of children of different ages.

Conclusions
To sum up, parenting plays a vital role in child development. There is an urgent need for a shorter and more reliable measure to evaluate different parenting styles and the effectiveness
of parental intervention programmes. The results of this study suggested that the proposed PS-7 had a better factor structure and psychometric properties than the original and other shortened versions of the PS. PS-7 also possesses good internal consistency and criterion validity, with the results being comparable to those for the full version of the PS. The seven-item version of the PS can provide a cost-effective method for assessing parenting practices and conducting epistemological surveys.

Supporting information

S1 Dataset.
(ZIP)

S1 Appendix. Factor structure of parenting scale and different shortened versions.
(DOCX)

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References

1. Belfer ML. Child and adolescent mental disorders: The magnitude of the problem across the globe. Journal of Child Psychology and Psychiatry. 2008; 49(3):226–36. https://doi.org/10.1111/j.1469-7610.2007.01855.x PMID: 18221350
2. Belsky J. The Determinants of Parenting: A Process Model. Child Development. 1984; 55(1):83–96. https://doi.org/10.1111/j.1467-8624.1984.tb00275.x PMID: 6705636
3. Morawska A, Winter L, Sanders MR. Parenting knowledge and its role in the prediction of dysfunctional parenting and disruptive child behaviour. Child: Care, Health and Development. 2009; 35(2):217–26.
4. Sanders MR. Triple P-Positive Parenting Program: Towards an Empirically Validated Multilevel Parenting and Family Support Strategy for the Prevention of Behavior and Emotional Problems in Children. Clinical Child and Family Psychology Review. 1999; 2(2):71–90. https://doi.org/10.1023/a:1021843613840 PMID: 11225933
5. Chua A. Battle hymn of the tiger mother. New York: New York: Penguin Press; 2011.
6. Kim SY, Wang YJ, Orozco-Lapray D, Shen YS, Murtuza M. Does "Tiger Parenting" Exist? Parenting Profiles of Chinese Americans and Adolescent Developmental Outcomes. Asian Am J Psychol. 2013; 4(1):7–18. https://doi.org/10.1037/a0030612 PMID: 23646228
7. Kelley ML, Tseng HM. Cultural-Differences in Child Rearing: A Comparison of Immigrant Chinese and Caucasian American Mothers. Journal of Cross-Cultural Psychology. 1992; 23(4):444–55.
8. Okagaki L, Frensch PA. Parenting and children’s school achievement: A multiethnic perspective. Am Educ Res J. 1998; 35(1):123–44.
9. Leung K, Lau S, Lam WL. Parenting styles and academic achievement: A cross-cultural study. Merrill-Palmer Q-J Dev Psychol. 1998; 44(2):157–72.
10. Chen XY, Dong Q, Zhou H. Authoritative and authoritarian parenting practices and social and school performance in Chinese children. International Journal of Behavioral Development. 1997; 21(4):855–73.
11. Chao RK. Beyond parental control and authoritarian parenting style: understanding Chinese parenting through the cultural notion of training. Child Development. 1994; 65(4):1111–9. https://doi.org/10.1111/j.1467-8624.1994.tb00806.x PMID: 7956468
12. Chao RK. Extending research on the consequences of parenting style for Chinese Americans and European Americans. Child Development. 2001; 72(6):1832–43. https://doi.org/10.1111/1467-8624.00381 PMID: 11768148
13. Wu P, Robinson CC, Yang C, Hart CH, Olsen SF, Porter CL, et al. Similarities and differences in mothers’ parenting of preschoolers in China and the United States. International Journal of Behavioral Development. 2002; 26(6):481–91.
14. Rohner RP, Khaleque A, Cournoyer DE. Parental acceptance-rejection: Theory, methods, cross-cultural evidence, and implications. Ethos. 2005; 33(3):299–334.
15. Arnold DS, O’Leary SG, Wolff LS, Acker MM. The Parenting Scale: A measure of dysfunctional parenting in discipline situations. Psychological Assessment. 1993; 5(2):137–44.
16. Klem S, Lohmann A, Möste T, Foran HM, Hahlweg K, Zenger M, et al. Development and Validation of a Parenting Scale Short Form (PS-8) in a Representative Population Sample. Journal of Child and Family Studies. 2019; 28(1):30–41.
17. Zhu AYF. Validating the Scale Measuring Dysfunctional Parenting with Hong Kong Adolescents. Child and Adolescent Social Work Journal. 2018; 35(5):489–98.
18. Prinzie P, Onghena P, Hellinckx W. Reexamining the parenting scale—Reliability, factor structure, and concurrent validity of a scale for assessing the discipline practices of mothers and fathers of elementary-school-aged children. European Journal of Psychological Assessment. 2007; 23(1):24–31.
19. Tahmouresi N, Schmiltz J, Bender C, Tuschens-Caffier B. The Impact of Culture on Parenting and Psychopathology in Children: A Comparative Study Between Iran and Germany. Iranian Journal of Psychiatry and Behavioral Sciences. 2017; 11(1):e4178.
20. Salari R, Terreros C, Sarkadi A. Parenting Scale: Which Version Should We Use? Journal of Psychopathology and Behavioral Assessment. 2012; 34(2):268–81.
21. Del Vecchio T, Jerusalim D, Terjesen MD. Psychometric characteristics of the Parenting Scale in a Vietnamese sample. International Journal of Psychology. 2017; 52(6):482–90. https://doi.org/10.1002/ijop.2242 PMID: 28644267
22. Brodard F, El Ghaziri N, Kounou KB, Zecca G. Validation de la version française d’une échelle évaluant les pratiques disciplinaires de parents d’enfants d’âge scolaire. Journal de Thérapie Comportementale et Cognitive. 2018; 28(3):114–22.
23. Itani T. The Japanese version of the Parenting Scale: Factor structure and psychometric properties. The Japanese Journal of Psychology. 2010; 81(5):446–52. https://doi.org/10.4992/jjpsy.81.446 PMID: 21226282
24. Suzuki K, Kita Y, Kaga M, Takehara K, Misago C, Inagaki M. The association between children’s Behavior and Parenting of caregivers: a longitudinal study in Japan. Front Public Health. 2016; 4:6. https://doi.org/10.3389/fpubh.2016.00006
25. Dumas JE, Amiga X, Bègle AM, Longoria Z. “When Will Your Program Be Available in Spanish?” Adapting an Early Parenting Intervention for Latino Families. Cognitive and Behavioral Practice. 2010; 17(2):176–87. https://doi.org/10.1016/j.cbpra.2010.01.004 PMID: 20077140
26. Karazsia BT, van Dulmen MMH, Wildman BG. Confirmatory Factor Analysis of Arnold et al.’s Parenting Scale Across Race, Age, and Sex. Journal of Child and Family Studies. 2008; 17(4):500–16.
27. Arney F, Rogers H, Baghurst P, Sawyer M, Prior M. The reliability and validity of the Parenting Scale for Australian mothers of preschool-aged children. Australian Journal of Psychology. 2008; 60(1):44–52.
28. Harvey E, Danforth JS, Ulaszek WR, Eberhardt TL. Validity of the parenting scale for parents of children with attention-deficit/hyperactivity disorder. Behaviour Research and Therapy. 2001; 39(6):731–43. https://doi.org/10.1016/s0005-7967(00)00352-8 PMID: 11400716

29. Bigras M, LaFreniere PJ, Dumas JE. Discriminant Validity of the Parent and Child Scales of the Parenting Stress Index. Early Education and Development. 1996; 7(2):167–78.

30. Rhoades KA, O’Leary SG. Factor structure and validity of the parenting scale. J Clin Child Adolesc Psychol. 2007; 36(2):137–46. https://doi.org/10.1037/0005-7967(98)00114-4 PMID: 9990744

31. Reitman D, Currier RO, Hupp SDA, Rhode PC, Murphy MA, O’Callaghan PM. Psychometric characteristics of the Parenting Scale in a head start population. J Clin Child Psychol. 2001; 30(4):514–24. https://doi.org/10.1207/S15374424JCCP3004_08 PMID: 11708239

32. Hu L-T, Bentler PM. Cutoff criteria for fit indexes in covariance structure analysis: Conventional criteria versus new alternatives. Structural Equation Modeling: A Multidisciplinary Journal. 1999; 6(1):1–55.

33. Kline RB. Principles and practice of structural equation modeling. 2 ed: New York: Guilford Press; 2005.

34. Hair JF. Multivariate data analysis. 7 ed: Upper Saddle River, NJ: Prentice Hall; 2010.

35. Irvine AB, Biglan A, Smolkowski K, Ary DV. The value of the Parenting Scale for measuring the discipline practices of parents of middle school children. Behaviour Research and Therapy. 1999; 37 (2):127–42. https://doi.org/10.1016/s0005-7967(98)00114-4 PMID: 9990744

36. Ford D, Chui WH. Where east meets west: Fieldwork instruction in Hong Kong and England. Asia Pac J Soc Work. 2001; 10(2):19–39.

37. Wong YLR. When East meets West—Nation, colony, and Hong Kong women’s subjectivities in gender and China development. Mod China. 2004; 30(2):259–92.

38. Beattie P, Bettache K, Chong KCY. Who is the neoliberal? Exploring neoliberal beliefs across East and West. J Soc Issues. 2019; 75(1):20–48.

39. Brislin RW. Back-Translation for Cross-Cultural Research. Journal of Cross-Cultural Psychology. 1970; 1(3):185–216.

40. Cha ES, Kim KH, Erlen JA. Translation of scales in cross-cultural research: issues and techniques. J Adv Nurs. 2007; 58(4):386–95. https://doi.org/10.1111/j.1365-2648.2007.04242.x PMID: 17442038

41. Loewenthal KM. An introduction to psychological tests and scales. 2 ed: Philadelphia, Pa.: Psychology Press; 2001.

42. Schel SHH, Bouman YHA, Vorstenbosch ECW, Bulten BH. Development of the forensic inpatient quality of life questionnaire: short version (FQL-SV). Quality of Life Research. 2017; 26(5):1153–61. https://doi.org/10.1007/s11136-016-1461-9 PMID: 27878427

43. MacKenzie MB, Kocovski NL, Blackie RA, Carrique LC, Fleming JE, Antony MM. Development of a Brief Version of the Social Anxiety—Acceptance and Action Questionnaire. Journal of Psychopathology and Behavioral Assessment. 2017; 39(2):342–54.

44. Markus A, Kokkinos CM. Development of a short form of the Greek Big Five Questionnaire for Children (GBFQ-C-SF): Validation among preadolescents. Personality and Individual Differences. 2017; 112:12–7.

45. Smith GT, McCarthy DM, Anderson KG. On the sins of short-form development. Psychological Assessment. 2000; 12(1):102–11. https://doi.org/10.1037/1040-3590.12.1.102 PMID: 10752369

46. Svedholm-Hakkinen AM, Lindeman M. Actively open-minded thinking: development of a shortened scale and disentangling attitudes towards knowledge and people. Think Reasoning. 2018; 24(1):21–40.

47. Chae D, Park Y. Development and Cross-Validation of the Short Form of the Cultural Competence Scale for Nurses. Asian Nurs Res. 2018; 12(1):69–76.

48. Zhang XT, Wang MC, He LN, Je L, Deng JX. The development and psychometric evaluation of the Chinese Big Five Personality Inventory-15. Plos One. 2019; 14(8):21.

49. Cattell RB. The Scree Test For The Number Of Factors. Multivariate Behavioral Research. 1966; 1 (2):245–76. https://doi.org/10.1207/s15327906mbr0102_10 PMID: 26828106

50. Fokkema M, Greiff S. How Performing PCA and CFA on the Same Data Equals Trouble Overfitting in the Assessment of Internal Structure and Some Editorial Thoughts on It. European Journal of Psychological Assessment. 2017; 33(6):399–402.

51. Babayak MA. What you see may not be what you get: A brief, nontechnical introduction to overfitting in regression-type models. Psychosom Med. 2004; 66(3):411–21. https://doi.org/10.1097/01.psy. 0000127692.23278.a9 PMID: 15184705
53. Jennrich RI, Sampson PF. Rotation for simple loadings. Psychometrika. 1966; 31(3):313–23. https://doi.org/10.1007/bf02289465 PMID: 5221128
54. Field AP. Discovering statistics using IBM SPSS statistics. 5th edition. ed. Field AP, editor: Los Angeles, California: SAGE Publications; 2018.
55. Cronbach LJ. Coefficient alpha and the internal structure of tests. Psychometrika. 1951; 16(3):297–334.
56. Zinbarg RE, Revelle W, Yovel I, Li W. Cronbach’s alpha, Revelle’s beta, and McDonald’s (omega H): Their relations with each other and two alternative conceptualizations of reliability. Psychometrika. 2005; 70(1):123–33.
57. Revelle W, Zinbarg RE. Coefficients Alpha, Beta, Omega, and the glb: Comments on Sijtsma. Psychometrika. 2009; 74(1):145–54.
58. McDonald RP. Test theory: a unified treatment: London: L. Erlbaum Associates; 1999.
59. Tabachnick BG. Using multivariate statistics. 6 ed. Fidell LS, editor: Boston: Pearson Education; 2013.
60. Brown TA. Confirmatory Factor Analysis for Applied Research, Second Edition: New York: Guilford Publications; 2014.
61. Jöreskog KG. A general approach to confirmatory maximum likelihood factor analysis. Psychometrika. 1969; 34(2):183–202.
62. DiStefano C, Morgan GB. A Comparison of Diagonal Weighted Least Squares Robust Estimation Techniques for Ordinal Data. Structural Equation Modeling: A Multidisciplinary Journal. 2014; 21(3):425–38.
63. Li C-H. Confirmatory factor analysis with ordinal data: Comparing robust maximum likelihood and diagonally weighted least squares. Behavior Research Methods. 2016; 48(3):936–49. https://doi.org/10.3758/s13428-015-0619-7 PMID: 26174714
64. Lionetti F, Kejser S, DellaGiulia A, Pastore M. Evidence of factorial validity of parental knowledge, control and solicitation, and adolescent disclosure scales: When the ordered nature of Likert scales matters. Frontiers in Psychology. 2016; 7, 941. https://doi.org/10.3389/fpsyg.2016.00941 PMID: 27445909
65. Fung S. Psychometric Evaluation of the Warwick-Edinburgh Mental Well-being Scale (WEMWBS) with Chinese University Students. Health and Quality of Life Outcomes. 2019; 17:46. https://doi.org/10.1186/s12955-019-1113-1 PMID: 30871563
66. Fung S. Cross-cultural validation of the Social Media Disorder scale. Psychol Res Behav Manag. 2019; 12:683–90. https://doi.org/10.2147/PRBM.S216788 PMID: 31695527
67. Browne MW, Cudeck R. Alternative ways of assessing model fit. In: Bollen KA, Long JS, editors. Testing structural equation models: Newbury Park: Sage; 1993. p. 136–62.
68. Schreiber JB, Nora A, Stage FK, Barlow EA, King J. Reporting Structural Equation Modeling and Confirmatory Factor Analysis Results: A Review. The Journal of Educational Research. 2006; 99(6):323–38.
69. Bentler PM, Bonett DG. Significance tests and goodness of fit in the analysis of covariance structures. Psychological Bulletin. 1980; 88(3):588–606.
70. Raine A, Dodge K, Loeber R, Gatze-Kopp L, Lynam D, Reynolds C, et al. The reactive–proactive aggression questionnaire: differential correlates of reactive and proactive aggression in adolescent boys. Aggressive Behavior. 2006; 32(2):159–71. https://doi.org/10.1002/ab.20115 PMID: 20798781
71. Fung AL-C, Raine A, Gao Y. Cross-Cultural Generalizability of the Reactive–Proactive Aggression Questionnaire (RPOQ). Journal of Personality Assessment. 2009; 91(5):473–9. https://doi.org/10.1080/00223890903088420 PMID: 19672753
72. Tuvblad C, Dhamija D, Berntsen L, Raine A, Liu J. Cross-Cultural Validation of the Reactive-Proactive Aggression Questionnaire (RPOQ) Using Four Large Samples from the US, Hong Kong, and China. Journal of Psychopathology and Behavioral Assessment. 2016; 38(1):48–55. https://doi.org/10.1007/s10862-015-9501-2 PMID: 27330246
73. Achenbach TM. Manual for the child behavior checklist and revised child behavior profile. Edelbrock CS, editor. Burlington, VT.: T.M. Achenbach; 1983.
74. Achenbach TM, Ruffle TM. The Child Behavior Checklist and Related Forms for Assessing Behavioral/Emotional Problems and Competencies. Pediatrics in Review. 2000; 21(8):265–71. https://doi.org/10.1542/pir.21-8-265 PMID: 10922023
75. Ivanova MY, Achenbach TM, Dumenci L, Rescorla LA, Almqvist F, Weintrob S, et al. Testing the 8-Syndrome Structure of the Child Behavior Checklist in 30 Societies. Clinical Child & Adolescent Psychology. 2007; 36(3):405–17.
76. Liu J, Cheng H, Leung PWL. The Application of the Preschool Child Behavior Checklist and the Caregiver–Teacher Report Form to Mainland Chinese Children: Syndrome Structure, Gender Differences,
77. Conners CK, Sitarenios G, Parker JDA, Epstein JN. The Revised Conners' Parent Rating Scale (CPRS-R): Factor Structure, Reliability, and Criterion Validity. Journal of Abnormal Child Psychology. 1998; 26(4):257–68. https://doi.org/10.1023/a:1022602400621 PMID: 9700518

78. Gau SS-F, Soong W-T, Chiu Y-N, Tsai W-C. Psychometric Properties of the Chinese Version of the Conners' Parent and Teacher Rating Scales-Revised: Short Form. Journal of Attention Disorders. 2006; 9(4):648–59. https://doi.org/10.1177/1087054705284241 PMID: 16648232

79. Robinson CC, Mandleco B, Olsen SF, Hart CH. Authoritative, Authoritarian, and Permissive Parenting Practices: Development of a New Measure. Psychological Reports. 1995; 77(3):819–30.

80. Foster SL, Cone JD. Validity issues in clinical assessment. Psychological Assessment. 1995; 7(3):248–60.

81. Rosseel Y. lavaan: An R Package for Structural Equation Modeling. Journal of Statistical Software. 2012; 48(2):36.

82. Hermida R. The Problem of Allowing Correlated Errors in Structural Equation Modeling: Concerns and Considerations. Computational Methods in Social Sciences. 2015; 3(1):5–17.

83. Wong KTW, Zheng V. Democratic Support and Cultural Values: An Empirical Study of Hong Kong and East Asian Societies. China. 2018; 16(2):111–32.

84. Cheng BS, Boer D, Chou LF, Huang MP, Yoneyama S, Shim D, et al. Paternalistic Leadership in Four East Asian Societies: Generalizability and Cultural Differences of the Triad Model. Journal of Cross-Cultural Psychology. 2014; 45(1):82–90.

85. Hwang KK. Face and favor—the chinese-power game. Am J Sociol. 1987; 92(4):944–74.

86. Offences against the Person Ordinance. Sect. Cap 212 (2017).