Remote Ischemic Preconditioning Fails to Benefit Pediatric Patients Undergoing Congenital Cardiac Surgery

A Meta-Analysis of Randomized Controlled Trials

Hong-Tao Tie, MSc, Ming-Zhu Luo, MSc, Zhen-Han Li, MSc, Qian Wang, MSc, Qing-Chen Wu, MD, Qiang Li, MD, and Min Zhang, MD

Abstract: Remote ischemic preconditioning (RIPC) has been proven to reduce the ischemia-reperfusion injury. However, its effect on children receiving congenital cardiac surgery (CCS) was inconsistent. We therefore performed the current meta-analysis of randomized controlled trials (RCTs) to comprehensively evaluate the effect of RIPC in pediatric patients undergoing CCS.

The outcomes included the duration of mechanical ventilation (MV), intensive care unit (ICU) length of stay, postoperative cardiac troponin (cTnI) level, hospital length of stay (HLOS), postoperative inotropic score, and mortality. Subgroup and sensitivity analysis were also performed as predesigned. The meta-analysis was performed with random-effects model despite of heterogeneity. Sensitivity and subgroup analysis were predesigned to identify the robustness of the pooled estimate.

Nine RCTs with 697 pediatric patients were included in the meta-analysis. Overall, RIPC failed to alter clinical outcomes of duration of MV (standard mean difference [SMD] = –0.03, 95% confidence interval [CI] = –0.23–0.17), ICU length of stay (SMD = –0.22, 95% CI = –0.47–0.04), or HLOS (SMD = –0.14, 95% CI = –0.55–0.26). Additionally, RIPC could not reduce postoperative cTnI (at 4–6 hours: SMD = –0.25, 95% CI = –0.73–0.23; P = 0.311; at 20–24 hours: SMD = 0.09, 95% CI = –0.51–0.68; P = 0.778) or postoperative inotropic score (at 4–6 hours: SMD = –0.19, 95% CI = –0.51–0.14; P = 0.264; at 24 hours: SMD = –0.15, 95% CI = –0.49–0.18; P = 0.365).

RIPC may have no beneficial effects in children undergoing CCS. However, this finding should be interpreted with caution because of heterogeneity and large-scale RCTs are still needed.

INTRODUCTION

Congenital heart disease (CHD) is the most common congenital anomaly with an incidence of approximately 0.9% and is a leading cause of mortality in neonates.1,2 With great advances in cardiothoracic surgery, 95% of neonates with CHD could survive to adulthood.3 Nevertheless, cardiopulmonary bypass was routinely used during congenital cardiac surgery (CCS)4 and could cause ischemia-reperfusion injury (IRI).5 IRI was associated with heart injuries, such as arrhythmia, myocardial stunning, low cardiac output, and perioperative myocardial infarction, especially for the immature newborn myocardium which was more vulnerable to apoptosis.6,7 Despite great progress in the underlying mechanism of myocardial IRI, numerous strategies to prevent heart injury are still disappointing.

Ischemic preconditioning (IPC) was reported to be a striking strategy that intermittent periods of myocardial ischemia could reduce the cardiac injury resulting from a subsequent prolonged period of ischemia.8 Consistently, the protection effect of IPC on ischemic injury was also confirmed in other organs in both animal and human.9–11 It was demonstrated that cardiac protection of IPC could be achieved by prevention against abnormal adenosine triphosphate production, mitochondrial swelling, and cell-membrane rupture.12,13 Additionally, IPC could induce expressions of nitric oxide synthase, cyclooxygenase 2, and antiapoptotic protein,14,15 to contribute to the cardiac protection. However, direct invasive IPC may be impractical and harmful in some clinical settings, and the access to temporarily occlude blood supply of the target organ is unavailable.16 Fortunately, transient ischemia in a remote tissue, namely remote ischemic preconditioning (RIPC), could also confer protection effect to target tissue from a subsequent lethal ischemia.17 RIPC shares the similar mechanism of IPC in cardiac protection, and it could also suppress the proinflammatory gene transcription and promote antiinflammation gene transcription.18

Three meta-analyses have demonstrated that RIPC could reduce the release of cardiac troponin (cTnI) in patients undergoing cardiac surgery.19–21 Specially, one of them, involving a total of 214 subjects, suggested that RIPC could provide cardiac protection in pediatric patients undergoing CCS.20 But in recent years, accumulating evidences of randomized controlled trials (RCTs)22–25 found that RIPC failed to show significant benefits on children undergoing CCS. Additionally,
our previous meta-analysis found that RIPC could only benefit the clinical outcome of intensive care unit (ICU) length of stay rather than the others. The primary outcome of cTnI was included in our inclusion criteria, and the exclusion of studies assessing other clinical outcomes might cause inadequate statistical power. Considering the statistical power and controversial effects of RIPC on children undergoing CCS, therefore we performed this meta-analysis of RCTs focusing on all potential clinical outcomes to give a comprehensive evaluation of RIPC in pediatric patients undergoing CCS.

METHODS

We conducted and reported the current meta-analysis in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-analysis statement (Additional file 1, http://links.lww.com/MD/A481). Since all analyses are on the basis of previous published studies, ethical approval and patient consent are not required.

Literature Search and Study Selection

PubMed, Embase, and Cochrane library were searched by 2 experienced authors independently, from the inception to July 2015 without any limitation. We combined MeSH and free-text terms to identify all potentially relevant studies with the following words: (“ischemic preconditioning” OR “myocardial ischemic preconditioning” OR “remote ischemic preconditioning” OR “limb ischemic preconditioning”) AND (“cardiovascular surgical procedures” OR “cardiac surgical procedures” OR “thoracic surgical procedures” OR “ventricular septal defects” OR “atrial septal defects” OR “cardiopulmonary bypass” OR “cardiac surgery” OR “heart surgery”) AND (child OR children OR infants OR infant OR newborn OR newborns OR neonate OR neonates). The bibliographies of retrieved studies were also screened for other relevant studies, and the process was performed repeatedly until no additional eligible studies were added. Two authors independently selected the studies by screening titles/abstracts, and then full texts were acquired for further screening. Disputes were finally adjudicated by the 3rd author.

Inclusion criteria: population, pediatric patients undergoing CCS; intervention, RIPC; control, sham operation; outcome, no limitation; study design, RCTs.

Data Extraction and Outcomes

Two investigators independently extracted data by using a standardized spreadsheet, and the following information were extracted: first author, publication year, country, sample size, patient baseline characteristic, surgery type, intervention of RIPC, intervention of control, start time of RIPC, and study design/Jadad score. Data were extrapolated from figures as needed, and Data.Graph.Digitizer.v2.24 was used to digitize graphs and plots according to the instruction. Although essential data were missing or uncertain, the corresponding authors would be contacted. All extracted data were rechecked by the 3rd author and disagreements were resolved by negotiation.

The outcomes were the duration of mechanical ventilation (MV), ICU length of stay, postoperative cTnI level, hospital length of stay (HLOS), postoperative inotropic score, and mortality.

Quality Assessment

Methodological quality of the included RCTs was independently assessed by 2 investigators using the modified Jadad scale. The scale includes 4 items: the generation of random sequences (0–2 points), blinding method (0–2 points), concealment of allocation (0–2 points), and withdrawal and drop-out (0–1 point). The quality score varies from 0 to 7 points, and a Jadad score >3 indicates a high-quality study. Discrepancies were resolved by discussion and adjudicated by the 3rd author.

Statistical Analysis

For all outcomes except mortality, standard mean differences (SMDs) with 95% confidence intervals (CIs) were used to calculate the estimated pooled difference between RIPC and control groups. Since no death was observed in each group in 6 included studies, the effect of RIPC on mortality was only qualitatively described. For the purpose of meta-analysis, continuous variables expressed as medians with ranges were converted to the means and the variances by an elementary inequalities and approximations. The random-effects model with DerSimonian and Laird weights was used in all analysis. Statistical heterogeneity was tested by the I² statistic and was considered to be low (I² less than 50%), moderate (I² between 50% and 75%), or high (I² more than 75%). Sensitivity and subgroup analysis were predefined to identify the robustness of the pooled estimate. The stratification factors for subgroups were children age (<1 vs >1 year), RIPC cycles (4 vs 3), and sample size (≥50 vs <50). Sensitivity analysis was performed via omitting 1 study and pooling the others in each turn. Additionally, sensitivity analysis was also conducted according to various inclusion criteria (children with ventricular septal defect [VSD], cyanosed children, RIPC in lower limb, RIPC after anesthesia, inflation pressure 15 mmHg > systolic arterial pressure [SAP], inflation pressure 30 mmHg > SAP, cTnI at postoperative 4 hours, cTnI at postoperative 6 hours, and cTnI at postoperative 24 hours). Publication bias was not evaluated because fewer than 10 publications were included. A 2-tailed P value less than 0.05 indicated a statistical significance. Stata 12.0 software (StataCorp, College Station, TX) was used for all statistical analysis.

RESULTS

Study Selection and Identification

A total of 185 publications were identified from the initial database search (PubMed [n = 49], Embase [n = 112], and Cochrane library [n = 24]). Of them, 48 were excluded owing to duplicate studies, 121 were excluded by screening titles/abstracts. Among the remaining 16 articles, 6 conference abstracts were excluded for the duplicate publications of the included full texts, which were also confirmed by the corresponding authors. Additionally, one was excluded for focusing on the intervention of IPC but not RIPC. Finally, 9 eligible RCTs meeting the inclusion criteria were included in the meta-analysis. The process of identification of eligible RCTs is presented in Figure 1.

Characteristics of Eligible Studies

Table 1 describes the main characteristics and Table S (Additional file 2, http://links.lww.com/MD/A481) presents the outcome data of the 9 included RCTs. These studies were published between 2006 and 2014 with a total of 697 pediatric patients from 6 countries. The samples ranged from 22 to 299. All except 3 trials involved pediatric patients with an average age of less than 1 year. The types of CHD were VSD, tetralogy of fallot (TOF), transposition of the
great arteries or hypoplastic left heart syndrome, and any form of CHD. The intervention of RIPC in all studies was induced by 3 or 4 cycles of 5-minute ischemia and 5-minute reperfusion using a blood-pressure cuff inflated to a pressure greater than the SAP, and the locations were around the lower limb in 8 trials and the upper limb in the remaining 1. The intervention of RIPC was carried out after or during anesthetic induction in 6 studies, 5 to 10 minutes before bypass in 1 study, and independently 24 and 1 hour before operation in the other 1. Among the included studies, eight reported the duration of MV, eight reported the ICU length of stay, seven reported the release of cTnI at postoperative 4 to 6 hours, and five reported inotropic score at postoperative 4 to 6 hours. All included RCTs were high-quality ones with Jadad scores.

**Duration of Mechanical Ventilation**

The pooled estimate of 8 studies with 592 children revealed that RIPC failed to shorten the duration of MV (SMD = 0.03, 95% CI = 0.23–0.17, I² = 17.7%, P = 0.758, Figure 2A). To explore the potential sources of heterogeneity and the impact of stratification factors on the pooled estimate, subgroup and sensitivity analysis were performed as predesigned. Table 2 shows the results of subgroup analysis and sensitivity analysis according to various inclusion criteria. Omitting one study and combining the remaining ones in each turn showed that the null association remained stable, with a range from −0.15 (95% CI = −0.38–0.08) to 0.06 (95% CI = 0.11–0.23).

**Intensive Care Unit Length of Stay**

Eight studies with 658 patients reported the outcome of ICU length of stay, and the combined estimate showed that RIPC could not significantly reduce the ICU length of stay (SMD = −0.22, 95% CI = −0.47–0.04, I² = 52.8%, P = 0.101, Figure 2B). Subgroup and sensitivity analysis are summarized in Table 3. Sensitivity analysis via omitting 1 study in each turn showed the results remained nonsignificant except for excluding the study by Cheung et al (SMD = −0.28, 95% CI = −0.53 to −0.03).

**Postoperative Cardiac Troponin**

Seven RCTs involving 293 pediatric patients were eligible for assessing the effect of RIPC on postoperative cTnI. Overall, RIPC could not decrease concentrations of cTnI at postoperative 4 to 6 hours (SMD = −0.25, 95% CI = −0.73–0.23, 0.25–0.73, and any form of CHD. The intervention of RIPC in all studies was induced by 3 or 4 cycles of 5-minute ischemia and 5-minute reperfusion using a blood-pressure cuff inflated to a pressure greater than the SAP, and the locations were around the lower limb in 8 trials and the upper limb in the remaining 1. The intervention of RIPC was carried out after or during anesthetic induction in 6 studies, 5 to 10 minutes before bypass in 1 study, and independently 24 and 1 hour before operation in the other 1. Among the included studies, eight reported the duration of MV, eight reported the ICU length of stay, seven reported the release of cTnI at postoperative 4 to 6 hours, and five reported inotropic score at postoperative 4 to 6 hours. All included RCTs were high-quality ones with Jadad scores.

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| Study ID            | Country     | No. of Patients | Population        | Surgery Type | Intervention | Start Time of RIPC | Location  | Jadad Score |
|---------------------|-------------|-----------------|-------------------|--------------|--------------|--------------------|-----------|-------------|
| Cheung et al 2006   | Canada      | 17/20           | 2.2 ± 3.4 years   | CHD/CPB      | 4 × 5 × 5 15 mmHg > SAP | Sham operation | Five minutes to ten minutes before initiation of bypass | Lower limb | 7           |
| Zhou et al 2010     | China       | 30/30           | 160.83 ± 58.39    | VSD/CPB      | 3 × 5 × 5 240 mmHg | Sham operation | Independently 24 hours and 1 hour before operation | Left upper arm | 7           |
| Luo et al 2011      | China       | 20/20           | 2.7 ± 0.9 years   | VSD/CPB      | 3 × 5 × 5 200 mmHg | Sham operation | Immediately after anesthetic induction | Lower limb | 7           |
| Lee et al 2012      | Korea       | 27/28           | 3.4 (R 0.6–12) months | VSD/CPB      | 4 × 5 × 5 30 mmHg > SAP | Sham operation | Approximately 10 minutes after induction of anesthesia | Lower limb | 7           |
| Pavione et al 2012  | Brazil      | 12/10           | 6.1 (R 1.4–13.9) months | CHD/CPB      | 4 × 5 × 5 15 mmHg > SAP | Sham operation | Twenty-four hours before cardiac surgery | Lower limb | 7           |
| Pedersen 2012       | Denmark     | 54/51           | 0.9 (0.2–5.0) years | CHD/CPB      | 4 × 5 × 5 40 mmHg > SAP | Sham operation | Immediately after anesthetic induction | Lower limb | 7           |
| Jones 2013          | Australia   | 20/19           | 5.5 ± 4.2 days    | TGA, HLHS/CPB | 4 × 5 × 5 15 mmHg > SAP | Sham operation | After induction of general anesthesia | Lower limb | 7           |
| Pepe et al 2013     | Australia   | 20/20           | 7.6 ± 3.2 months  | TOF/CPB      | 4 × 5 × 5 30 mmHg > SAP | Sham operation | Immediately after the induction of anesthesia | Lower limb | 7           |
| McGrinlde 2014      | Canada      | 148/151         | 2.2 (0.5 to 5.9) years | CHD/CPB      | 4 × 5 × 5 15 mmHg > SAP | Sham operation | During induction of anesthesia | Lower limb | 7           |

CHD = congenital heart disease, Con = control, HLHS = hypoplastic left heart syndrome, RCT = randomized controlled trial, RIPC = remote ischemic preconditioning, SAP = systolic arterial pressure, TGA = transposition of the great arteries, TOF = tetralogy of Fallot, VSD = ventricular septal defect. 4 × 5 × 5 15 mmHg > SAP: 4 cycles of 5-minute ischemia and 5-minute reperfusion using a blood-pressure cuff inflated to a pressure 15 mmHg greater than the systolic arterial pressure.
1^2 = 75.2%, \( P_H < 0.001 \), Figure 3A) or at postoperative 20 to 24 hours (SMD 0.09, 95% CI 0.51–0.68, \( I^2 = 83.7% \), \( P_H < 0.001 \); \( P = 0.778 \), Figure 3B). Table 4 summarizes the results of subgroup and sensitivity analysis. Further exclusion of any single study could not significantly alter the results with a range from −0.39 95% CI (−0.86–0.07) to −0.05 95% CI (−0.36–0.26) for cTnI at postoperative 4 to 6 hours and from −0.13 95% CI (−0.66–0.41) to 0.31 95% CI (−0.18–0.80) for cTnI at postoperative 20 to 24 hours.

**Hospital Length of Stay**

Five studies22,25,34–36 with 506 patients were included for the outcome of HLOS. RIPC failed to substantially alter HLOS (SMD −0.14, 95% CI −0.55–0.26, \( I^2 = 71.9% \), \( P_H = 0.007 \); \( P = 0.493 \), Figure 2C). Subgroup analysis and sensitivity analysis were not performed for outcomes of HLOS and postoperative inotropic score because of the limited amount of studies.

**Postoperative Inotropic Score**

Four studies23,25,33,37 with 192 reported the outcome of inotropic score at postoperative 4 to 6 hours and five22,25,33,36,37 with 232 at postoperative 20 to 24 hours. Overall, the pooled estimate revealed that RIPC intervention was not associated with a reduction in the amount of hemodynamic support at postoperative 4 to 6 hours (SMD −0.19, 95% CI −0.51–0.14, \( I^2 = 23.1% \), \( P_H = 0.272 \); \( P = 0.264 \), Figure 4A) or postoperative 24 hours (SMD −0.15, 95% CI −0.49–0.18, \( I^2 = 38.2% \), \( P_H = 0.167 \); \( P = 0.365 \), Figure 4B).

**DISCUSSION**

**Main Findings**

Our meta-analysis suggested RIPC failed to reduce postoperative cTnI, indicating RIPC may have no benefit for cardiac protection in children undergoing CCS. The results remained stable and robust in sensitivity and subgroup analysis. Additionally, RIPC could not reduce the amount of hemodynamic support or HLOS. Moreover, overall results demonstrated that RIPC failed to shorten the duration of MV or ICU length of stay; however, the results was not consistent in subgroup or sensitivity analysis.

**Comparison With Previous Studies**

Our finding was inconsistent with the previous meta-analysis,20 which, involving a total of 214 subjects, suggested that RIPC could provide cardiac protection in pediatric patients undergoing CCS (postoperative cTnI: SMD −0.75, 95%
CI − 1.05 to −0.46). However, substantial heterogeneity among studies were observed in both our study ($I^2 > 75\%$) and the previous one ($I^2 = 88\%$), indicating that the RIPC effect was more different from each other than random error could explain.

With adding 4 new RCTs of 79 patients, the controversial results might be explained by the increased statistical power. Additionally, time-point measurement data were used for cardiac protection analysis in our study, while area under the curve

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**TABLE 2.** Subgroup Analysis and Sensitivity Analysis According to Various Inclusion Criteria for the Duration of MV

| Subgroup/Sensitivity Analysis | No. of Trials | No. of Patients | SMD (95% CI) | $I^2$, % | $P_H$ | $P$ Value |
|------------------------------|--------------|----------------|-------------|---------|-------|--------|
| Total                        | 8            | 592            | −0.03 (−0.23−0.17) | 17.8     | 0.284 | 0.755  |
| Subgroup analysis            |              |                |             |         |       |        |
| Age                          |              |                |             |         |       |        |
| >1 year                      | 3            | 376            | 0.05 (−0.30−0.40) | 40.8    | 0.184 | 0.791  |
| <1 year                      | 5            | 216            | −0.15 (−0.42−0.11) | 0       | 0.591 | 0.260  |
| RIPC cycles                  |              |                |             |         |       |        |
| Four                         | 6            | 492            | −0.01 (−0.24−0.23) | 20.6    | 0.272 | 0.970  |
| Three                        | 2            | 100            | −0.17 (−0.60−0.27) | 17.6    | 0.271 | 0.459  |
| Sample size                  | <50           | 5              | −0.20 (−0.50−0.09) | 0       | 0.448 | 0.178  |
| ≥50                          | 3            | 414            | 0.12 (−0.07−0.31) | 0       | 0.407 | 0.227  |
| Sensitivity analysis         |              |                |             |         |       |        |
| Children with VSD            | 3            | 155            | −0.17 (−0.48−0.15) | 0       | 0.305 | 0.544  |
| Cyanosed children             | 2            | 79             | −0.43 (−0.87−0.02) | 0       | 0.913 | 0.062  |
| RIPC in lower limb           | 7            | 532            | −0.07 (−0.31−0.18) | 28.1    | 0.204 | 0.593  |
| RIPC after anesthesia         | 4            | 174            | −0.35 (−0.65−0.05) | 0       | 0.944 | 0.023  |
| Pressure 15 mmHg > SAP       | 4            | 397            | 0.15 (−0.05−0.35) | 0       | 0.613 | 0.136  |
| Pressure 30 mmHg > SAP       | 2            | 95             | −0.32 (−0.73−0.08) | 0       | 0.418 | 0.119  |

MV = mechanical ventilation, SMD = standard mean difference, $P_H = P$ for heterogeneity, RIPC = remote ischemic preconditioning, VSD = ventricular septal defect, SAP = systolic arterial pressure.

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**TABLE 3.** Subgroup Analysis and Sensitivity Analysis According to Various Inclusion Criteria for ICU Length of Stay

| Subgroup/Sensitivity Analysis | No. of Trials | No. of Patients | SMD (95% CI) | $I^2$, % | $P_H$ | $P$ Value |
|------------------------------|--------------|----------------|-------------|---------|-------|--------|
| Total                        | 8            | 658            | −0.22 (−0.47−0.04) | 52.8     | 0.038 | 0.101  |
| Subgroup analysis            |              |                |             |         |       |        |
| Age                          |              |                |             |         |       |        |
| >1 year                      | 3            | 376            | 0.01 (−0.31−0.33) | 31.6    | 0.232 | 0.945  |
| <1 year                      | 5            | 282            | −0.38 (−0.69−0.08) | 34.2    | 0.194 | 0.014  |
| RIPC cycles                  |              |                |             |         |       |        |
| Four                         | 6            | 558            | −0.23 (−0.57−0.11) | 64.9    | 0.014 | 0.179  |
| Three                        | 2            | 100            | −0.16 (−0.55−0.23) | 0       | 0.439 | 0.428  |
| Sample size                  | <50           | 7              | −0.16 (−0.70−0.37) | 59.1    | 0.062 | 0.556  |
| ≥50                          | 4            | 519            | −0.24 (−0.55−0.07) | 59.9    | 0.058 | 0.134  |
| Sensitivity analysis         |              |                |             |         |       |        |
| Children with VSD            | 3            | 155            | −0.36 (−0.78−0.06) | 41       | 0.183 | 0.089  |
| RIPC in lower limb           | 7            | 598            | −0.25 (−0.54−0.05) | 58.9    | 0.024 | 0.105  |
| RIPC after anesthesia         | 4            | 240            | −0.51 (−0.77−−0.25) | 0       | 0.568 | <0.001 |
| Pressure 15 mmHg > SAP       | 3            | 358            | 0.05 (−0.16−0.26) | 0       | 0.461 | 0.634  |
| Pressure 30 mmHg > SAP       | 2            | 95             | −0.75 (−1.16−0.33) | 0       | 0.940 | <0.001 |

ICU = intensive care unit, SMD = standard mean difference, $P_H = P$ for heterogeneity, RIPC = remote ischemic preconditioning, VSD = ventricular septal defect, SAP = systolic arterial pressure.

Pressure refers to inflation pressure.
Surgery were almost old people and usually accompanied with other diseases, such as diabetes mellitus, hypertension, coronary heart disease, and chronic obstructive pulmonary disease. However, our finding was contradicted with our previous research which found that RIPC could only benefit ICU length of stay and we attributed the null association with other outcomes to inadequate statistical power. However, the only beneficial effect of RIPC on ICU length of stay vanished when increasing the statistical power by adding 2 studies. Therefore, we could conclude that the phenomenon that no beneficial effect of RIPC on ICU length of stay was more prone to be a verdict than happened by chance. Inconsistency existed in results of the duration of MV and ICU length of stay in the subgroup and sensitivity analysis. However, considering limitations of observational analysis and decreased statistical power, the inconsistent results should be interpreted cautiously. Seven included trials reported mortality among children undergoing CCS. However, since 6 of them reported no death in each group, quantitative analysis could not be performed. All included studies consistently demonstrated that RIPC could not alter mortality in pediatric patients after CCS, suggesting that no association between RIPC and mortality existed.

In adult patients, a mass of evidence suggested that RIPC could reduce the release of cTnI while undergoing CABG and valve replacement. The results were consistent with 2 meta-analyses despite substantial heterogeneity. Why the promising effect was not observed in pediatric patients? The contradictory effect of RIPC between adults and children might be explained by the following aspects: First, children’s heart is immature and differs from adults. Adults suffer more from hypoxia. Third, the long duration of ischemia during cardiac surgery might overwhelm the protective effect of RIPC because the protective effect of RIPC might be masked or confounded by hypoxia. Third, the long duration of ischemia during cardiac surgery might overwhelm the protective effect of RIPC; consequently, diverse surgery time in various CHD might be involved in the controversy. Fourth, the interval between RIPC and the initiation of cardiopulmonary bypass might be another cause for the null effect, which was confirmed by our sensitivity analysis for the duration of MV and ICU length of stay. Finally, small sample size and substantial heterogeneity might also contribute to the controversy.

**Clinical Implication and Further Research**

Despite well feasibility and tolerance, RIPC should not be recommended to use in children undergoing CCS according to our findings. The different intervention of RIPC, anesthetic regimens, surgery type, cyanosed or acyanotic children, and inflation pressure could contribute to the inconsistency of the results and substantial heterogeneity, thus these confounding factors should be considered in further researches. With the advances in understanding the mechanism of RIPC in cardiac surgery, several pharmacological agents have been developed to simulate IPC intervention and some have shown a promising effect in animal or human. The pharmacological preconditioning might be another attractive way to trigger the similar beneficial effect of RIPC in pediatric patients to adult patients, and inevitably a number of basic and clinical researches are needed.
**TABLE 4. Subgroup Analysis and Sensitivity Analysis According to Various Inclusion Criteria for cTnI**

| Subgroup/Sensitivity Analysis | No. of Trials | No. of Patients | SMD (95%CI) | $I^2$ % | $P_H$ | $P$ Value | SMD (95%CI) | $I^2$ % | $P_H$ | $P$ Value |
|-------------------------------|--------------|----------------|-------------|--------|-------|-----------|-------------|--------|-------|-----------|
| **Total**<sup>22–25,33,36,37</sup> | 7            | 293            | -0.25 (-0.73–0.23) | 75.2   | <0.001 | 0.311     | 0.09 (-0.51–0.68) | 83.7   | <0.001 | 0.778     |
| **Subgroup analysis**         |              |                |             |        |       |           |             |        |       |           |
| Age                           |              |                |             |        |       |           |             |        |       |           |
| >1 year<sup>36,37</sup>       | 2            | 77             | -1.05 (-2.23–0.12) | 82.7   | 0.016 | 0.079     | -0.87 (-1.92–0.18) | 79.4   | 0.027 | 0.106     |
| <1 year<sup>22–25,33</sup>    | 5            | 216            | 0.02 (-0.31–0.35) | 31.7   | 0.210 | 0.889     | 0.44 (-0.10–0.98) | 73     | 0.005 | 0.109     |
| **RIPC cycles**               |              |                |             |        |       |           |             |        |       |           |
| Four<sup>2–2.5,37</sup>       | 5            | 193            | -0.01 (-0.40–0.38) | 44.9   | 0.123 | 0.963     | 0.39 (-0.23–1.01) | 77     | 0.002 | 0.218     |
| Three<sup>33,36</sup>         | 2            | 100            | -0.90 (-2.35–0.56) | 90.8   | 0.001 | 0.227     | -0.67 (-2.08–0.73) | 90.6   | 0.001 | 0.347     |
| **Sample size**               |              |                |             |        |       |           |             |        |       |           |
| <50<sup>22,24,25,36,37</sup>  | 5            | 178            | -0.25 (-1.01–0.50) | 83.4   | <0.001 | 0.511     | 0.16 (-0.78–1.11) | 88.9   | <0.001 | 0.738     |
| ≥50<sup>31,33</sup>           | 2            | 115            | -0.24 (-0.60–0.13) | 0      | 0.744 | 0.209     | -0.06 (-0.43–0.30) | 0      | 0.627 | 0.732     |
| **Sensitivity analysis**       |              |                |             |        |       |           |             |        |       |           |
| Children with VSD<sup>23,33,36</sup> | 3            | 155            | -0.67 (-1.49–0.14) | 83.3   | 0.003 | 0.107     | -0.48 (-1.27–0.31) | 82.6   | 0.003 | 0.232     |
| Cyanosed children<sup>24,25</sup> | 2            | 79             | 0.29 (-0.39–0.97) | 56.4   | 0.130 | 0.402     | 0.82 (-0.22–1.85) | 79.4   | 0.028 | 0.121     |
| RIPC in lower limb<sup>2–2.5,36,37</sup> | 6            | 233            | -0.26 (-0.85–0.33) | 79.3   | <0.001 | 0.385     | 0.10 (-0.64–0.84) | 86.4   | <0.001 | 0.787     |
| RIPC after anesthesia<sup>23–25,36</sup> | 4            | 174            | -0.33 (-1.18–0.52) | 86.6   | <0.001 | 0.449     | 0.02 (-1.00–1.04) | 90.5   | <0.001 | 0.967     |
| Pressure 15 mmHg > SAP<sup>22,24,37</sup> | 3            | 98             | 0.14 (-0.55–0.83) | 64.4   | 0.060 | 0.691     | 0.64 (-0.45–1.73) | 84.6   | 0.001 | 0.248     |
| Pressure 30 mmHg > SAP<sup>23,25</sup> | 2            | 95             | -0.19 (-0.60–0.21) | 0      | 0.550 | 0.346     | 0.04 (-0.41–0.49) | 18.2   | 0.269 | 0.856     |
| cTnI at postop 4 hours<sup>22,33</sup> | 2            | 82             | -0.06 (-0.50–0.37) | 0      | 0.391 | 0.777     | –            | –      | –      | –         |
| cTnI at postop 6 hours<sup>23–25,36,37</sup> | 5            | 211            | -0.35 (-1.02–0.32) | 82.3   | <0.001 | 0.303     | –            | –      | –      | –         |
| cTnI at postop 24 hours<sup>22–25,33,37</sup> | 6            | 253            | –            | –      | –      | –         | 0.31 (-0.18–0.80) | 72.5   | 0.003 | 0.213     |

cTnI = cardiac troponin, postop = postoperative, SMD = standard mean difference, $P_H = P$ for heterogeneity, VSD = ventricular septal defect, RIPC = remote ischemic preconditioning, SAP = systolic arterial pressure.

<sup>Pressure refers to inflation pressure.</sup>
Fourth, some other interesting clinical outcomes, such as left ventricular ejection fraction, HLOS, or mortality, suggesting that RIPC may not provide significant benefits in terms of these endpoints. However, larger-scale RCTs are still needed to further evaluate the potential clinical implications of RIPC.

In summary, RIPC fails to reduce the duration of MV, ICU length of stay, postoperative cTnI, amount of hemodynamic support, and HLOS, or mortality, suggesting that RIPC may not benefit pediatric patients undergoing CCS. However, our findings should be interpreted with caution because of the heterogeneity and hence large-scale RCTs are still needed.

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LIMITATIONS

Several limitations should be considered in our meta-analysis. First, the original data of cTnI in several studies were estimated from the figures, which would increase some additional errors. To reduce the error to the greatest extent, the data were estimated by 2 reviewers independently and rechecked by the 3rd investigator. Second, considerable heterogeneity was observed. In order to explore the potential source of heterogeneity and the stability of pooled effects, subgroup and sensitivity analysis were conducted. Although heterogeneity still existed, the pooled effects were consistent. Third, the anesthetic regimens have been estimated by 2 reviewers independently and rechecked by the 3rd investigator. Second, considerable heterogeneity was observed. To reduce the error to the greatest extent, the data were estimated from the figures, which would increase some additional errors. Therefore, the impact of the pooled effect was not explored because of unavailable data. Fourth, some other interesting clinical outcomes, such as left ventricular ejection fraction, cardiac function, and renal function, were not investigated because of sparse and inconsistent reportings among studies.

CONCLUSIONS

In summary, RIPC fails to reduce the duration of MV, ICU length of stay, postoperative cTnI, amount of hemodynamic support, HLOS, or mortality, suggesting that RIPC may not benefit pediatric patients undergoing CCS. However, our findings should be interpreted with caution because of the heterogeneity, and hence large-scale RCTs are still needed.
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