Aberrant levels of cortical myelin distinguish individuals with depressive disorders from healthy controls

David A.A. Baranger\textsuperscript{a, *}, Yaroslav O. Halchenko\textsuperscript{b}, Skye Satz\textsuperscript{a}, Rachel Ragozzino\textsuperscript{a}, Satish Iyengar\textsuperscript{c}, Holly A. Swartz\textsuperscript{b}, Anna Manelis\textsuperscript{a, *}

\textsuperscript{a} Department of Psychiatry, Western Psychiatric Institute and Clinic, University of Pittsburgh Medical Center, University of Pittsburgh, Pittsburgh, PA, USA
\textsuperscript{b} Department of Psychological and Brain Sciences, Dartmouth College, NH, USA
\textsuperscript{c} Department of Statistics, University of Pittsburgh, Pittsburgh, PA, USA

\textbf{ARTICLE INFO}

\textbf{Keywords:}
Depression
Cortical myelin
MRI
Machine learning
T1w/T2w ratio
Elastic net
LDA

\textbf{ABSTRACT}

The association between depressive disorders and measures reflecting myelin content is underexplored, despite growing evidence of associations with white matter tract integrity. We characterized the T1w/T2w ratio using the Glasser atlas in 39 UD and 47 HC participants (ages = 19–44, 75% female). A logistic elastic net regularized regression with nested cross-validation and a subsequent linear discriminant analysis conducted on held-out samples were used to select brain regions and classify patients vs. healthy controls (HC). True-label model performance was compared against permuted-label model performance. The T1w/T2w ratio distinguished patients from HC with 68% accuracy (p < 0.001; sensitivity = 63.8%, specificity = 71.5%). Brain regions contributing to this classification performance were located in the orbitofrontal cortex, anterior cingulate, extended visual, and auditory cortices, and showed statistically significant differences in the T1w/T2w ratio for patients vs. HC. As the T1w/T2w ratio is thought to characterize cortical myelin, patterns of cortical myelin in these regions may be a biomarker distinguishing individuals with depressive disorders from HC.

\textbf{1. Introduction}

Depressive disorders, also collectively referred as unipolar depression (UD), is a leading cause of disability worldwide (World Health Organisation, 2017), with an economic burden of $210 billion dollars in the United States alone (Greenberg et al., 2015). Despite its impact, treatments for the disorder remain ineffective for many patients (Cuijpers et al., 2020). Thus, there is a pressing need to understand the neurobiological etiology of UD to facilitate the development of improved treatments and prevention strategies.

Unipolar depression is characterized by dysfunctional affective and cognitive processing (Gotlib and Joormann, 2010), including reduced executive functioning (Snyder, 2013), biased emotional processing (Suslow et al., 2020), and impaired reward processing (Halakoon et al., 2020). Correspondingly, individuals with UD show aberrant activation during tasks which recruit these processes, including activation in the striatum, hippocampus, amygdala, orbitofrontal cortex, prefrontal cortex, insula, cingulate, and occipital cortex (Hamilton et al., 2012; Ng et al., 2019; Stuhrmann et al., 2011; Wang et al., 2015). In addition, a growing body of literature has reported structural abnormalities associated with depression in many of these same regions described above, both in grey matter (Ho et al., 2020; Schmaal et al., 2017, 2016) and in white matter (Manelis et al., 2021; Shen et al., 2017; van Velzen et al., 2020) Meta-analyses of diffusion weighted imaging (DWI) studies have repeatedly found evidence for lower fractional anisotropy (FA) in depressed populations (Liao et al., 2013; Murphy and Frodl, 2011; Wise et al., 2016). More recently, studies using very large samples (i.e., the UK Biobank), as well as meta-analyses combining both published and unpublished data (i.e., the ENIGMA consortium), have observed widespread and replicable reductions in FA (Shen et al., 2017; van Velzen et al., 2020). Notably, white matter integrity in identified regions has been shown to correlate with the cognitive processes disrupted in depression, including processing speed (Chopra et al., 2018; Penke et al., 2010), emotion regulation (Eden et al., 2015; Welton et al., 2020), and reward learning (de Boer et al., 2020).

Emerging evidence suggests that cortical myelin may be impacted in individuals with UD and that it may partially mediate some of the cognitive processes that are impaired in affected individuals. For
example, studies of post mortem brain tissue from donors with UD have observed reduced myelination, a reduction in the number of oligodendrocytes and other glia (cells whose functions include generating and maintaining myelin (Simons and Nave, 2016)), and reduced expression of oligodendrocyte lineage genes (Boda, 2021; Tham et al., 2011). A study of individuals with treatment resistant depression revealed a reduced magnetization transfer ratio (MTR), which is thought to reflect lower myelin levels, in the cingulate cortex and insula (Zhang et al., 2009). A recent study using R1 (1/T1) as a measure of myelination observed reduced whole-brain myelin, but no significant difference in cortical myelin in a handful of a priori bilateral regions between individuals with depression and healthy controls (HC) (Sacchet and Gotlib, 2017).

Developments in magnetic resonance imaging (MRI) methodology permit the examination of cortical myelin via the T1w/T2w ratio (Glasser and van Essen, 2011; Shafee et al., 2015). While the T1w/T2w ratio does not solely reflect cortical myelin (Fukutomi et al., 2018; Hagiwara et al., 2018; Tardif et al., 2017; Yasuno et al., 2017), it has been shown to correlate with several white matter metrics (Ganzetti et al., 2014; Glasser and van Essen, 2011; Ritchie et al., 2018; Shams et al., 2019), indicating that it is sensitive to cortical myelin. Studies in population-based samples using this metric have found that lower myelin in the cingulate, orbitofrontal cortex, and middle temporal cortex correlated with poor sleep quality (Toschi et al., 2020), lower frontal-pole myelin and greater myelin in the occipital cortex correlated with neuroticism (Toschi and Passamonti, 2019), and lower myelin in the motor and higher myelin in the insular, cingulate, prefrontal, and superior parietal cortices correlated with trait anxiety (Norbom et al., 2019). While poor sleep, high neuroticism, and trait anxiety might represent concurrent symptoms of depression, prior studies have not systematically examined cortical myelin in participants with depression as compared to HC.

The goals of the present study were (1) to ascertain whether the T1w/T2w ratio is predictive of UD, and (2) to characterize the brain regions that are predictive of case/control status. Based on the prior studies mentioned above, we hypothesized that the T1w/T2w ratio would distinguish individuals with UD from HC and that these differences will be especially pronounced in the prefrontal cortical (PFC), cingulate, parietal and occipital regions that support reward and emotional processing, which are dysregulated in UD (Dalili et al., 2015; Ng et al., 2019).

2. Methods

2.1. Participants

The study was approved by the University of Pittsburgh Institutional Review Board. Participants were recruited from the community, universities, and counseling and medical centers. They gave written informed consent, were right-handed, fluent in English, and were matched on age and sex. Individuals with unipolar depression (UD) met DSM-5 criteria for major depressive or persistent depressive disorders. Healthy controls (HC) had no personal or family history of psychiatric disorders. Exclusion criteria included a history of head injury, metal in the body, pregnancy, claustrophobia, neurodevelopmental disorders, systemic medical illness, premorbid IQ < 85 per the National Adult Reading Test (Nelson, 1982), current alcohol/drug abuse, Young Mania Rating Scale scores > 10 (YMRS (Young et al., 1978)) at scan, or meeting criteria for any psychotic-spectrum disorder. Data were drawn from an ongoing longitudinal study that includes neuroimaging sessions at baseline and 6-month follow-up and clinical evaluations at baseline, 6-months, and 12-months. The present report includes baseline data available for 55 HC and 50 UD. Participants were excluded from analyses due to (1) previously undetected brain abnormalities of potential clinical relevance: 2 UD, (2) diagnosis conversion during the course of the study: 1 HC was diagnosed with major depressive disorder, and 1 UD was diagnosed with bipolar disorder; (3) scanner or movement-related artifacts in MRI data (4 HC, 7 UD), and (4) poor-quality myelin maps (see Section 2.4.2.1 Subject-level processing: 3 HC, 1 UD). The final sample included 47 HC and 39 UD.

2.2. Clinical assessment

All diagnoses were made by a trained clinician and confirmed by a psychiatrist according to DSM-5 criteria using SCID-5 (First, 2015). Additional information collected included illness onset and duration, number of current episodes, comorbid psychiatric disorders, current depression symptoms using the Hamilton Rating Scale for Depression (HRSD-25) (Hamilton, 1960), current mania symptoms using the Young Mania Rating Scale (YMRS) (Young et al., 1978), and lifetime depression and hypo/mania spectrum symptomatology using the Mood Spectrum Self Report (MOODS-SR) (Dell’Osso et al., 2002). A total psychotropic medication load was calculated for each participant with UD, with greater numbers and doses of medications corresponding to a greater medication load (Hassel et al., 2008; Manelis et al., 2016).

2.3. Neuroimaging data acquisition

The neuroimaging data were collected at the University of Pittsburgh/UPMC Magnetic Resonance Research Center using a 3 T Siemens Prisma scanner with a 64-channel receiver head coil and named according to the ReproNim convention (Visconti di Oleggio Castello et al., 2020). The DICOM images were converted to BIDS dataset using heudiconv (Halchenko et al., 2019) and dem2nix (Li et al., 2016). High-resolution T1w images were collected using the MP_RAGE sequence with TR = 2400 ms, resolution = 0.8x0.8x0.8 mm, 208 slices, FOV = 256, TE = 2.22 ms, flip angle = 8°. High-resolution T2w images were collected using TR = 3200 ms, resolution = 0.8x0.8x0.8 mm, 208 slices, FOV = 256, TE = 563 ms. Field maps were collected in the AP and PA directions using the spinecho sequence (TR = 8000, resolution = 2x2x2mm, FOV = 210, TE = 66 ms, flip angle = 90°, 72 slices).

2.4. Data analyses

2.4.1. Clinical data analysis

HC and UD groups were compared on demographic and clinical variables using t-tests and chi-square tests. All analyses were conducted in R (https://www.r-project.org/).

2.4.2. Neuroimaging data processing

2.4.2.1. Subject-level preprocessing. Data quality was examined using mriqc version 0.15.1 (Esteban et al., 2017) and visually inspected (Supplemental Methods). Each participant’s cortical myelin was characterized with the T1w/T2w ratio (Glasser et al., 2016, 2013; Glasser and van Essen, 2011) using the FreeSurfer, FreeSurfer, and PostFreeSurfer minimal preprocessing pipelines for the human connectome project (Glasser et al., 2013). Workbench v1.4.2 and HCPpipelines-4.1.3 were installed system-wide on a workstation with GNU/Linux Debian 10 operating system. The spin echo field maps collected in AP and PA phase encoding directions were used for bias field correction in PreFreeSurfer. Registration to standard space was achieved via MSMulc (Robinson et al., 2018) in PostFreeSurfer. If FreeSurfer images and myelin maps had artifacts and gross errors (e.g., large regions of apparent low myelin in the occipital cortex due to the transverse sinus interfering with accurate identification of the pial surface), the data were removed from analyses. The resulting myelin maps were parcellated in Workbench using the 360 region Glasser Atlas (Glasser et al., 2016). Further data quality assurance resulted in the removal of 11 outlier parcels (Supplemental Methods), leaving 349 parcels.
2.4.3. Neuroimaging data analysis

2.4.3.1. Elastic net and linear discriminant analysis. Neuroimaging data decoding studies can capitalize on complex relationships between variables, but their large numbers can present a challenge in deriving a generalizable model. The elastic net approach has emerged as a flexible tool for use with neuroimaging data (Jollans et al., 2019) as it is able to reduce the influence of overly large coefficients and reducing the number of variables while generating multivariate models predictive of complex behaviors (Acuff et al., 2019; Bertocci et al., 2016; Manelis et al., 2020). Elastic net is a regularized regression which combines lasso and ridge regression (i.e., L1- and L2- norm regularization) (Zou and Hastie, 2005). Ridge regression penalizes overly large coefficients, while lasso regression removes variables with small coefficients.

We used logistic elastic net regularized regression (Friedman et al., 2010) to select variables (brain parcels) that were most predictive of case/control status. Elastic net has two parameters: alpha ($\alpha$) controls the balance between the ridge and lasso regularizations, and lambda ($\lambda$) controls the strength of regularization. To provide equal contribution of each penalty to the loss function, we used $\alpha = 0.5$. To avoid model overfitting and bias, we implemented nested cross-validation to identify the optimal $\lambda$ parameter. A linear discriminant analysis (LDA) (Venables and Ripley, 2002) model was subsequently trained using selected variables to make out-of-sample prediction on held-out participants. This strategy is illustrated in Fig. 1A and described in detail in Supplemental Methods. For each repetition of the nested cross-validation loop, two participants (1 UD and 1 HC) were held out. The rest of the sample was used to identify the optimal $\lambda$ parameter which were then used to fit the elastic net model and select variables whose myelin levels were predictive of UD/HC status. These variables were then used to train an LDA model, which was tested on held-out participants. Results (model fit, variable selection, prediction accuracy) were evaluated against results from a permuted-label analysis (Fig. 1B and Supplemental Methods). To contribute to the ongoing discussion regarding best-practices for nested cross-validation with neuroimaging data sets (Vabalas et al., 2019; Varoquaux, 2018), post-hoc analyses systematically varied cross-validation parameters (i.e., the size of the held-out sample and the number of folds used to fit the elastic net model) and examined their impact on prediction accuracy and variable selection.

2.4.3.2. Post-hoc analyses

2.4.3.2.1. The relationship between true-label and permuted-label sample demographics. To ensure that the permuted results were not due to changes in the internal structure of the permuted samples in terms of demographic variables (i.e., age, sex and IQ), we compared the age, sex and IQ values in the permuted samples with that in the true-label samples.

Fig. 1. Diagram of analysis steps. Conceptual depiction of analysis steps including: (1) a unique pair of one UD and one HC participant is held-out; (2) an elastic net regression is used to select variables; (3) the retained variables are used an LDA model predicting case/control status; (4) the LDA model is tested on the held-out sample; (5) this process is repeated for each of the $n = 1833$ pairs of subjects; (6) for each held-out pair, the training procedure is repeated with 100 unique permutations.
Table 1 - Demographic and clinical characteristics.

|                  | HC (mean/sd or count/percent) | UD (mean/sd or count/percent) | t-test or chi-squared test HC vs. UD |
|------------------|-------------------------------|-----------------------------|-----------------------------------|
| N                | 47 (54.7%)                    | 39 (45.3%)                  |                                   |
| Gender (number females) | 36 (76.6%)                    | 29 (74.6%)                  |                                   |
| UD diagnoses (MDD/PDD) | na                           | 26/13                       |                                   |
| Age (years)      | 28.55 (6.15)                  | 29.07 (6.88)                | t(84) = -0.37, p = 0.71           |
| IQ (NART)        | 106.72 (6.24)                 | 110.2 (7.49)                | t(84) = -2.35, p = 0.02           |
| Lifetime onset (year of age) | na                           | 15.05 (5.02)                |                                   |
| Lifetime episodes of depression | na                          | 3.25 (1.5)                  |                                   |
| Current depression severity (HRSD-25) | 1.74 (2.16)                  | 12.69 (6.77)                | t(84) = -10.48, p < 0.01          |
| Lifetime depression (MOODS-SR) | 2.15 (2.27)                  | 18.41 (4.36)                | t(84) = -22.21, p < 0.01          |
| Number taking Antidepressants | na                          | 22 (56.4%)                  |                                   |
| Number taking Mood stabilizers | na                          | 3 (7.7%)                    |                                   |
| Number taking Antipsychotics | na                          | 1 (2.6%)                    |                                   |
| Number taking Benzodiazepines | na                          | 5 (12.8%)                   |                                   |
| Number taking Stimulants | na                          | 3 (7.7 %)                   |                                   |
| Mean number of psychotropic medications | na                          | 1.05 (1.12)                 |                                   |
| Mean total medication load | na                          | 1.33 (1.53)                 |                                   |
| Number with comorbid diagnoses | na                          | 27 (69.2%)                  |                                   |

Note: Comparison of demographic and clinical characteristics between healthy controls (HC) and individuals with Unipolar Depression (UD). Tests were run as t-tests or chi-squared tests, as specified. Bold = p < 0.05. na = not applicable.

2.4.3.2.2. Association of LDA accuracy and cortical myelin with demographic and clinical characteristics. To further characterize the parcels selected by the logistic elastic net regression, we compared cortical myelin in UD vs. HC while controlling for age, sex, and IQ. To assess the potential influence of confounding variables on model fit, we tested whether demographic or clinical characteristics were predictive of classification accuracy. Regressions tested whether the joint effect of group, the variable in question, or their interaction, was associated with participant-wise accuracy (Dinga et al., 2020). Additional models tested whether clinical characteristics of UD participants were associated with model accuracy. To assess the influence of confounding variables on variable selection, regressions similarly tested the association of demographic and clinical variables with cortical myelin, controlling for age, sex, and IQ. Results for each variable were separately corrected for multiple comparisons using false discovery rate (FDR).

2.4.3.3. Exploratory analysis of a HC participant who was diagnosed with major depressive disorder 12 months from the baseline scan. One participant entered the study as a HC but was diagnosed with major depressive disorder sometime between 6 and 12 months after study onset. At the study visit at 12 months the participant had mild depressive symptoms. While this participant was excluded from all primary analyses described above, exploratory analyses investigated whether the myelin was predictive of the participant’s conversion from HC to UD. This analysis used the primary 86 participants and the variables selected in primary analyses (cortical myelin in 33 parcels and IQ) to train an LDA model. UD/HC status was then predicted at study onset and at the 6-month follow-up (both time points were prior to conversion).

3. Results

3.1. Sample demographics

Individuals with UD did not differ by age or sex but had higher IQ and current and lifetime depression severity compared to HC (Table 1).

3.2. UD vs. HC LDA nested cross-validation classification accuracy

The T1w/T2w ratio and IQ distinguished UD from HC in subjects held-out during cross-validation, with an average accuracy of 68% (Fig. 2A; sensitivity (UD): 63.8%, specificity (HC): 71.5%). The mean participant-wise accuracy across 84 classification loops ranged from 0% to 100%, and 84% of participants were classified either quite accurately (>80% n = 53) or inaccurately (<20% n = 19). When demographic variables were excluded from the LDA classification, nested cross-validation achieved 69% accuracy (UD: 65.2%, HC: 73%). While the mean classification accuracy was higher in the HC group, this difference was not statistically significant (t = 0.88, p = 0.38). Notably, post-hoc analyses found that accuracy, sensitivity, and specificity were always significantly above chance across all cross-validation schemes examined (i.e., varying the size of the held-out sample and the number of folds used to fit the elastic net model) (Supplemental Fig. 3).

In permutation analyses, no variable was selected in 66.9% of 183,300 models. Within each participant, the proportion of models that did not select any variables ranged from 65.4% to 68.7%. Excluding instances when no variables were selected, the average participant-wise LDA accuracy in permutation analyses ranged from 41.2% to 59.2%, with an average accuracy of 50.5% (i.e., chance level). The test of whether age, sex and IQ in permuted samples differed from that in the sample with true labels showed that the demographics of the permuted groups differed (p < 0.05, uncorrected) from the true-label groups only in 1.7% of cases.

3.3. Elastic net variable selection

True-label nested cross-validated elastic net models predicting diagnostic status (UD vs. HC) with 349 parcels, age, sex and IQ, selected 90 myelin parcels (Supplemental Table 1) and IQ in at least one model (Fig. 2B). Each model selected between 9 and 68 variables, with a median of 17 variables. Permutated-label nested cross-validated elastic net models selected all predictor variables at least once, but no variable was selected by >3% of the models. In addition, 66.9% of models with permuted labels did not select any variable at all. Within permuted-label models where at least one variable was selected, the number of selected variables ranged from 1 to 113, with a median of 17 variables selected.

Given that the variable selection frequency in the models with permuted labels likely represents noise, we applied the criterion of the median + 3.5*IQR across all permuted variables (3.77%) as the cutoff value to separate the potential noise variables from ‘signal’ variables in the nested cross-validated elastic net with true labels. This latter analysis identified 33 out of 90 parcels, plus IQ, across the true-label nested cross-validated elastic net models that were above the cutoff line (Table 2, Fig. 2B). These 33 parcels included multiple regions in the orbitofrontal cortex, insula, cingulate, and frontal operculum, as well as regions in the auditory and visual cortices (Supplemental Fig. 3). No model retained the age and sex variables. Notably, post-hoc analyses which tested different cross-validation schemes found that these 34 variables were selected across all parameter variations (Supplemental Fig. 4).
Post-hoc analyses tested the association of diagnostic status with the T1w/T2w ratio in the 33 selected parcels (Table 2, Fig. 3). These parcels include both regions where UD participants have a lower mean T1w/T2w ratio than HC and regions where UD participants have a greater mean T1w/T2w ratio, as well as regions where the two groups do not differ. After FDR-correction for multiple comparisons, fourteen of the parcels showed evidence of significant differences between UD and HC participants (Table 2). An additional 8 parcels showed nominally significant differences between UD and HC participants (Table 2; p < 0.05 uncorrected). Parcels that showed a greater absolute mean difference between the groups were selected more frequently in the nested cross-validation analysis ($r = 0.7$, $p = 4 \times 10^{-6}$).

![Fig. 2. Classification accuracy and variable selection. A) Average subject-wise nested cross-validation classification accuracy in both healthy controls (HC) and participants with unipolar depression (UD). B) Variable selection frequency (percent of models where a given parcel was retained) in the true data (light blue) and permutations (green). The n = 90 parcels which were retained at least once are shown. IQ, the only demographic variable selected, is labeled. Solid black lines represent the median of the variable selection frequency for all n = 350 variables + 3.5 × the interquartile range (IQR). See Supplemental Table 2 for the full list of variable selection frequencies. (For interpretation of the references to colour in this figure legend, the reader is referred to the web version of this article.)

Table 2
The comparison of the T1w/T2w ratio in selected parcels in individuals with UD vs. HC.

| Parcel | t   | p    | p-FDR | Selection Frequency | Extended Description                  |
|--------|-----|------|-------|---------------------|---------------------------------------|
| Left 11 | −3.056 | 0.003 | 0.032 | 100.00 | BA 11 (orbital and polar frontal) |
| Left 7PC | 3.081 | 0.003 | 0.032 | 100.00 | BA 7 (superior parietal cortex) |
| Left MST | 2.892 | 0.005 | 0.032 | 100.00 | Medial superior temporal area |
| Left p10p | −2.536 | 0.013 | 0.038 | 100.00 | BA 10 (frontopolar margin of orbital prefrontal cortex) |
| Right FOP2 | −2.574 | 0.012 | 0.038 | 100.00 | Frontal Opercular area 2 (posterior opercular cortex) |
| Right FFC | −2.534 | 0.013 | 0.038 | 100.00 | Fusiform face complex |
| Left STSdp | −2.036 | 0.045 | 0.071 | 99.95 | Auditory association cortex |
| Right TE2p | 2.928 | 0.004 | 0.032 | 99.89 | Lateral temporal complex |
| Left Lbelt | −2.713 | 0.008 | 0.038 | 99.35 | Lateral belt complex (auditory) |
| Left a24pr | −3.003 | 0.004 | 0.032 | 99.29 | Ventral anterior cingulate cortex. |
| Right 7AL | −2.512 | 0.014 | 0.038 | 99.18 | BA 7 (superior parietal cortex) |
| Left d23ab | −2.157 | 0.034 | 0.060 | 95.53 | Ventral posterior cingulate cortex |
| Left 6a | −2.291 | 0.025 | 0.054 | 95.25 | BA 6 (premotor subdivisions) |
| Left 24dv | −2.556 | 0.012 | 0.038 | 93.36 | Ventral anterior cingulate cortex. |
| Right V3A | 1.795 | 0.076 | 0.097 | 47.52 | Visual area V3A |
| Left V4c | −1.650 | 0.103 | 0.041 | 39.44 | Visual area V4c |
| Right Lbelt | −2.122 | 0.037 | 0.061 | 33.99 | Lateral belt complex (auditory) |
| Left p32 | 2.179 | 0.032 | 0.060 | 33.66 | BA 32 (pregenual anterior cingulate) |
| Right PH | 2.152 | 0.034 | 0.060 | 22.97 | Posterior temporal visual region |
| Right FOS2 | 1.921 | 0.058 | 0.082 | 19.97 | Parieto-occipital sulcus area 2 |
| Right FOP3 | −1.816 | 0.073 | 0.096 | 19.31 | Frontal Opercular area 1 (posterior opercular cortex) |
| Right Rl | 1.559 | 0.123 | 0.235 | 18.22 | Retrolingular cortex |
| Left 47m | −1.735 | 0.086 | 0.102 | 15.77 | BA 47 (orbital part of inferior frontal gyrus) |
| Left c32 | −2.442 | 0.017 | 0.040 | 12.22 | BA 32 (subgenual anterior cingulate) |
| Left MT | 2.176 | 0.033 | 0.060 | 11.95 | Middle temporal area |
| Left RI | −2.520 | 0.014 | 0.038 | 11.78 | Retrolingular cortex |
| Right OP1 | −1.237 | 0.220 | 0.220 | 8.69 | Parietal operculum (secondary somatosensory cortex) |
| Left PeEc | −1.429 | 0.157 | 0.162 | 6.98 | Perirolateral ectorhinal cortex |
| Left FOP4 | −2.489 | 0.015 | 0.038 | 6.66 | Frontal Opercular area 4 (posterior opercular cortex) |
| Right RSC | −2.099 | 0.048 | 0.072 | 6.22 | Retrolingular complex |
| Right 52 | 1.493 | 0.139 | 0.148 | 6.06 | BA 52 (parasplenial) |
| Right 7PM | 1.913 | 0.059 | 0.082 | 5.18 | BA 7 (superior medial parietal cortex) |
| Right 5m | −1.777 | 0.079 | 0.097 | 3.98 | BA5 (paracentral lobule) |

The association of UD with the T1w/T2w ratio in the n = 33 selected variables. Tests were run as linear regressions, testing the association of case/control status with the T1w/T2w ratio, controlling for age, sex, and IQ. Positive t-values indicate greater T1w/T2w ratio (and, consequently, greater cortical myelin level) in UD vs. HC. We report both uncorrected and FDR-corrected p-values.
The 33 selected parcels included 13 parcels in regions important for executive functioning and cognitive control, located in the anterior cingulate (left a24pr, left 24dv, left p32, and left s32), orbitofrontal cortex (left 11 l, left p10p, and left 47 m), posterior cingulate (right RSC, and left d23ab), frontal operculum (right FOP1 and left FOP4), and parietal cortex (right POS2 and right 7Pm). UD was largely associated with a reduced T1w/T2w ratio in these regions, except in one region of the anterior cingulate (left p32), and the two regions in the parietal cortex (right POS2 and right 7Pm), where it was associated with an increased T1w/T2w ratio. An additional 19 of the identified parcels play roles in visual, somatomotor, and auditory processing. Visual processing regions included extrastriate regions (right V3A, left V4t, and left MT), ventral stream regions (right FFC and right TE2p), and dorsal stream regions (left MST, right PH, left 7PC, and right 7AL). Auditory processing regions included regions in the auditory cortex (left LBelt, right LBelt, left RI, and right RI) and regions implicated in language (left STSdp, left FOP2, and right 52). Somatomotor regions included regions implicated in somatosensation (right 5 m and right OP1) and a region in the pre-motor cortex (left 6a). UD was largely associated with a reduced T1w/T2w ratio in auditory and somatomotor regions, except for two auditory regions (right 52 and right RI) where it was associated with an increased T1w/T2w ratio. In contrast, participants with UD had a greater T1w/T2w ratio in 6 visual regions and had a reduced T1w/T2w ratio in 3 regions (left V4t, left FFC, and left 7AL).

### 3.4. Post-hoc analyses

#### 3.4.1. Association of the T1w/T2w ratio with demographic and clinical variables

The association of the T1w/T2w ratio in the 33 selected parcels with demographic and clinical variables (see Table 1) was not statistically significant (Supplemental Table 4). Across all participants neither sex, IQ, nor MOODS-SR score were correlated with cortical myelin in selected parcels, nor did they interact with group (HC/UD) to predict cortical myelin. There was evidence for associations between age and the T1w/T2w ratio in two regions (p-fdr < 0.05), driven by a positive correlation in right LBelt, and a negative correlation in left a24pr. In individuals with UD, no clinical variable was correlated with the T1w/T2w ratio in selected parcels. However, taking antidepressants was nominally associated (p < 0.05 uncorrected) with an increased T1w/T2w ratio in right OP1, left 24dv, and left p32, and a decreased T1w/T2w ratio in right LBelt (Supplemental Table 4, Supplemental Figure 5).

#### 3.4.2. Association of LDA accuracy with demographic and clinical variables

Analyses found that no demographic or clinical variable was predictive of classification accuracy (Table 3). Similarly, within the UD participant group, no clinical or medication variable was associated with classification accuracy (Table 3).

#### 3.4.3. Exploratory analysis of classification in a participant who converted from HC to UD

The LDA trained on the whole sample of 86 participants, with IQ and the 33 parcels identified in the previous analyses as predictors classified this participant as ‘UD’ both times: 12 months and 6 months before illness onset.

### Table 3

| Sample  | Variable            | F    | p    |
|---------|---------------------|------|------|
| All     | IQ × Group          | 0.5  | 0.68 |
| All     | Age × Group         | 1.65 | 0.18 |
| All     | Sex × Group         | 2.44 | 0.07 |
| All     | HRSBD-25 × Group    | 0.73 | 0.54 |
| All     | MOODS-SR × Group    | 1.13 | 0.34 |
| UD Only | Antidepressants     | 0.17 | 0.69 |
| UD Only | Age of illness onset| 0.39 | 0.53 |
| UD Only | Illness duration    | 0.43 | 0.52 |
| UD Only | Lifetime episodes of depression | 1.67 | 0.2 |
| UD Only | Number of comorbid diagnoses | 0.13 | 0.72 |
| UD Only | UD Diagnosis (MDD/PDD) | 2.04 | 0.16 |
| UD Only | Medication load     | 0.46 | 0.5  |

Joint regressions were run as Group × Variable + Group + Variable, predicting classification accuracy. No test was significant (all ps > 0.05). Medications other than antidepressants were not examined, as too few participants were taking them (Table 1).
4. Discussion

The T1w/T2w ratio distinguished healthy control participants (HC) from individuals with unipolar depression (UD) with 68% accuracy. The elastic net regression selected regions implicated in executive function and reward processing, as well as those involved in visual, auditory, and somatomotor processing. The pattern of T1w/T2w ratio distinguishing UD from HC was associated with a lower T1w/T2w ratio in some regions and a higher T1w/T2w ratio in other regions. As the T1w/T2w ratio partially reflects cortical myelin, these results demonstrate that aberrant levels of cortical myelin may be a biomarker of depressive disorders.

The 33 selected parcels included 13 parcels in regions important for executive functioning, cognitive control, and reward processing, including the anterior and posterior cingulate, orbitofrontal cortex, frontal operculum, and parietal cortex. There is abundant evidence that these processes are disrupted in depression (Ng et al., 2019; Snyder, 2013). The present results converge with mounting evidence of disrupted myelination (Zhang et al., 2009), thickness (Schmaal et al., 2017), connectivity (Weiskopf et al., 2006), and activation (Ng et al., 2019) of these regions in those with depression. In addition to regions traditionally reported in neuroimaging studies of depression, we found that parcels located along the dorsal and ventral visual pathways, somatosensory and auditory processing regions and regions implicated in language also contributed to UD vs. HC classification. While depression is not typically considered a disorder of dysfunctional sensory processing, some studies report visual (Grange and Rydon-Grange, 2020) and auditory (Zweerings et al., 2019) processing deficits in depression. In addition, somatosensory symptoms (e.g., motor retardation) are among the diagnostic criteria for this disorder and have been well described (Simon et al., 1999). Our results add to a growing body of literature documenting associations of depression with altered structure (Schmaal et al., 2017) and functional connectivity (Chen et al., 2019; Korgaonkar et al., 2020) in sensory regions. Broadly, our results contribute to the consensus that the neurobiology of regions that play important roles in cognitive processes and information transfer are disrupted in depression, which may contribute to the etiology of the disorder.

It is notable that in contrast to studies of the major white matter tracts, where depression is associated with lower integrity (van Velzen et al., 2020), we observed both a decreased and an increased T1w/T2w ratio in individuals with UD relative to HC. Additionally, the anterior cingulate, the superior parietal cortex, the retrosininal cortex, and visual processing regions included a combination of proximal regions with T1w/T2w ratio decreases and increases. While the relationship between the T1w/T2w ratio and functional activation or connectivity remains under-explored, these observations suggest that cortical myelin imbalance, rather than a uniform reduction, could drive some of the observed functional differences in depression, such as disrupted network integration (Chen et al., 2019; Zheng et al., 2018) and patterns of both hypo- and hyper- connectivity (Kaiser et al., 2015; Yu et al., 2019).

Intriguingly, while classification accuracy was bimodal, in permutations participants were classified at chance (50%). This suggests that participants who were consistently misclassified differed from those who were correctly classified in their demographic or clinical features. However, post-hoc analyses did not reveal any associations between classification accuracy and demographic or clinical variables, thus suggesting that misclassification could reflect the presence of environmental or genetic risk factors that are not captured by clinical measures. These results suggest that the myelination of a region may be creating predisposition to depression at the brain level. As a preliminary exploration of this idea, we explored the T1w/T2w ratio in a participant who experienced the onset of UD during the course of the study and was not used in primary analyses. This participant completed two MRI scans: at baseline and 6 months follow-up with both time points occurring prior to UD onset. Remarkably, this participant was classified as UD on both scans, despite not yet meeting criteria for a UD diagnosis at the time of scan. This preliminary result suggests that the T1w/T2w ratio in frontal, sensorimotor and extended visual cortices may be a biological risk marker predictive of UD diagnosis in the future. Further longitudinal studies are needed to test this hypothesis.

While the present results demonstrate that the T1w/T2w ratio is disrupted in UD, the cause of the disruption remains unknown. Post-hoc analyses suggest the T1w/T2w ratio is not correlated with medication use, lifetime depression severity, or illness duration. However, the T1w/T2w ratio may reflect other risk factors for UD. For instance, sleep disturbance, which is a well-established risk factor and symptom of UD (Franzen and Buyssse, 2008), was recently shown to correlate with cortical myelin in several of the same regions found in the present study, including the cingulate and middle temporal cortex (Tochchi et al., 2020). Stress is also associated with reduced white matter integrity in depression (Ho et al., 2017; van Velzen et al., 2020; Ziegler et al., 2020). After stress, remyelination can occur (Bonnefils et al., 2019) thus resulting into altered patterns of myelination across the cortex (Orthmann-Murphy et al., 2020). This ‘remyelination’ hypothesis could potentially explain our observations of both decreased and increased cortical myelin in unipolar depression.

Limitations of this work includes the need to replicate our findings in an independent sample. To partially address this limitation, our analyses used robust machine learning methods involving model testing using held-out samples. The major strength of this approach is that it helps to reduce model bias, which occurs when the same participants are used to train and test a model (Davatzikos, 2019). It has been suggested that model performance with nested cross-validation is close to the accuracy that would be achieved on fully independent data (Varma and Simon, 2006). The second limitation concerns measurement noise due to susceptibility artifacts. Several regions with documented relevance to depression, including the bilateral hippocampus, entorhinal cortex, and posterior orbitofrontal cortex complex, were not included in the present analyses, as these regions showed an excess of between-person variability (see Supplemental Methods). Future research should explore the association of the T1w/T2w ratio in these regions with unipolar depression.

In summary, the T1w/T2w ratio can distinguish participants with UD from HC, even when clinical and demographic variables are not included in analyses. Regions that were most important for this classification include several that play key roles in reward and emotion processing as well as a host of regions important for sensory processing. This result highlights that the association of UD with sensory processing bears further investigation. Notably, UD was associated with both a decreased and an increased T1w/T2w ratio, suggesting that observations of reduced integrity of major white matter tracts in UD may not fully extend to the cortex. These results suggest that cortical myelin holds promise as a biomarker of unipolar depression and may be an early predictor of risk for this disorder.

CRedit authorship contribution statement

David A.A. Baranger: Formal analysis, Data curation, Visualization, Writing – original draft. Yaroslav O. Halchenko: Data curation, Formal analysis, Writing - review & editing. Skye Satz: Investigation, Data curation, Visualization, Project administration, Writing - review & editing. Rachel Ragozzino: Investigation, Data curation, Project administration, Writing - review & editing. Satish Iyengar: Formal analysis, Writing - review & editing. Holly A. Swartz: Project administration, Writing - review & editing. Anna Manelis: Conceptualization, Funding acquisition, Project administration, Supervision, Investigation, Data curation, Formal analysis, Writing – original draft.

Acknowledgments

Funding Acknowledgements: This work was supported by a grant from the National Institute of Health R01MH114870 to A.M., and Y.O.H was supported by P41EB019936 to the Center for Reproducible
We also thank Dr. Mary L. Phillips for fruitful discussions of the study.

**Data availability**

All neuroimaging data are available in BIDS format in OpenNeuro repository under the accession number ds003653 at https://doi.org/10.18112/openneuro.ds003653.v1.0.0.

**References**

Acuff, H.E., Versace, A., Bertocci, M.A., Ladouceur, C.D., Hanford, L.C., Manelis, A., Acuff, H.E., Versace, A., Axelsson, J., Riklund, K., Nyberg, L., Backman, L., Guitart-Masip, M., 2019. Cortical and subcortical white matter integrity and dopamine D1 receptor availability predict age differences in prefrontal value signaling during reward processing. Cerebral Cortex. doi:10.1093/cercor/bha104.

Davatzikos, C., 2019. Machine learning in neuroimaging: Progress and challenges. NeuroImage 197, 652–656. https://doi.org/10.1016/j.neuroimage.2018.10.003.

Dell’OssO, L., Armani, A., Rucci, P., Frank, E., Fagiolini, A., Corretti, C., Shear, M.K., Grochocinski, V.J., Maser, J.D., Endicott, J., Castano, G.B., 2002. Measuring mood spectrum: Comparison of interview (SCINFOODS) and self-report (MOODS-SR) instruments. Compr. Psychiatry 43 (1), 69–73. https://doi.org/10.1016/S0010-9295(02)00301.

Dinga, R., Schmaa, L., Penninx, B.W.J.H., Veltman, D.J., Marquand, A.F., 2020. Controlling for Effects of Confounding Variables on Machine Learning Predictions. bioRxiv.

Eden, A.S., Schreiber, J., Anwander, A., Kremer, K., Laeger, I., Zwanzger, P., Zilles, K., 2003. The intra-subject variability of the human thalamus. Neuroimage 19 (1), 106–117. https://doi.org/10.1016/S1053-8119(02)00210-2.

Findling, R.L., Phillips, M.L., 2016. Predicting clinical outcome from reward circuitry function and white matter structure in behaviorally and emotionally dysregulated youth. Mol. Psychiatry 21 (9), 1194–1201. https://doi.org/10.1038/mp.2015.127.

Fukutomi, H., Glasser, M.F., Zhang, H., Autio, J.A., Coalson, T.S., Okada, T., Togashi, K., Van Essen, D.C., Hayashi, T., 2018. Neurite imaging reveals microstructural variations in human cerebral cortical gray matter. NeuroImage 182, 488–499. https://doi.org/10.1016/j.neuroimage.2018.02.017.

Ganzetti, M., Wenderoth, N., Mantini, D., 2014. Whole brain myelin mapping using T1- and T2-weighted MR imaging data. Front. Hum. Neurosci. 8, 671. https://doi.org/10.3389/fnhum.2014.00671.

Glasser, M.F., Coalson, T.S., Robinson, E.C., Hacker, C.D., Harwell, J., Yacoub, E., Ugurbil, K., Andersson, J., Beckmann, C.F., Jenkinson, M., Smith, S.M., Van Essen, D.C., 2016. A multi-modal parcellation of human cerebral cortex. Nature 536 (7615), 171-178. https://doi.org/10.1038/nature19035.

Glasser, M.F., Sotiropoulos, S.N., Wilson, J.A., Coalson, T.S., Fischl, B., Andersson, J.L., Xu, J., Jbabdi, S., Webster, M., Polimeni, J.R., Van Essen, D.C., Jenkinson, M., 2013. The minimal preprocessing pipelines for the Human Connectome Project. NeuroImage 80, 105–124. https://doi.org/10.1016/j.neuroimage.2013.04.127.

Glasser, M.F., Van Essen, D.C., 2011. Mapping human cortical areas in vivo based on myelin content as revealed by T1- and T2-weighted MRI. J. Neurosci. 31 (32), 11862–11867. https://doi.org/10.1523/jneurosci.146211.2011.

Grazulis, P., Stringaris, A., Wolpert, M., 2020. Treatment outcomes for depression: challenges and opportunities. The Lancet Psychiatry 7 (11), 925–927. https://doi.org/10.1016/S2215-0366(20)30036-5.
Venables, W.N., Ripley, B.D., 2002. Modern Applied Statistics with S, Fourth edi, ed. Springer.

Visconti di Oleggio Castello, M., Dobson, J.E., Sackett, T., Kodiweera, C., Hashy, J. V., Goncalves, M., Ghosh, S., Halchenko, Y.O., 2020. ReproNim/repronin 0.6.0. doi: 10.5281/ZENODO.3625000.

Wang, X.-L., Du, M.-Y., Chen, T.-L., Chen, Z.-Q., Huang, X.-Q., Luo, Y.a., Zhao, Y.-J., Kumar, P., Gong, Q.-Y., 2015. Neural correlates during working memory processing in major depressive disorder. Prog. Neuro-Psychopharmacol. Biol. Psychiatry 56, 101–108. https://doi.org/10.1016/j.pnpbp.2014.08.011.

Weiskopf, N., Hutton, C., Josephs, O., Deichmann, R., 2006. Optimal EPI parameters for reduction of susceptibility-induced BOLD sensitivity losses: A whole-brain analysis at 3 T and 1.5 T. NeuroImage 33 (2), 493-504. https://doi.org/10.1016/j.neuroimage.2006.07.029.

Welton, T., Indja, B.E., Maller, J.J., Fanning, J.P., Valleye, M.P., Grieve, S.M., 2016. Voxel-based meta-Analytical evidence of structural disconnectivity in major depression and bipolar disorder. Biol. Psychiatry 79 (4), 293–302. https://doi.org/10.1016/j.biopsych.2015.03.004.

World Health Organisation, 2017. Depression and other common mental disorders: global health estimates. World Health Organization.

Yasuno, F., Kazui, H., Morita, N., Kajimoto, K., Ihara, M., Taguchi, A., Yamamoto, A., Matsuoka, K., Takahashi, M., Nakagawara, J., Iida, H., Kishimoto, T., Nagatuka, K., 2017. Use of T1-weighted/T2-weighted magnetic resonance ratio to elucidate changes due to amyloid β accumulation in cognitively normal subjects. NeuroImage: Clinical 13, 209-214. https://doi.org/10.1016/j.nicl.2016.11.029.

Young, R.C., Biggs, J.T., Ziegler, V.E., Meyer, D.A., 1978. A rating scale for mania: Reliability, validity and sensitivity. Br. J. Psychiatry 133 (5), 429–435. https://doi.org/10.1192/bjp.133.5.429.

Yu, M., Linn, K.A., Shinhara, R.T., Oathes, D.J., Cook, P.A., Duprat, R., Moore, T.M., Oquendo, M.A., Phillips, M.L., McInnis, M., Fava, M., Trivedi, M.H., McGrath, P., Parsey, R., Weissman, M.M., Shelene, Y.L., 2019. Childhood trauma history is linked to abnormal brain connectivity in major depression. PNAS 116 (17), 8062-8090. https://doi.org/10.1073/pnas.1900801116.

Zhang, T.J., Wu, Q.Z., Huang, X.Q., Sun, X.L., Zou, K., Lui, S., Liu, F., Hu, J.M., Kuang, W. H., Li, D.M., Li, F., Chen, H.F., Chan, R.C.K., Mechelli, A., Gong, Q.Y., 2009. Magnetization transfer imaging reveals the brain deficit in patients with treatment-refractory depression. J. Affect. Disord. 117 (3), 157–161. https://doi.org/10.1016/j.jad.2009.01.003.

Zou, H., Hastie, T., 2005. Regularization and variable selection via the elastic net. J. Royal Statistical Soc. Series B: Statistical Methodology. 67 (2), 301–320. https://doi.org/10.1111/rssb.2005.67.issue-210.1111/j.1467-9868.2005.00503.x.

Zweerings, J., Zvyagintsev, M., Turetsky, B.L., Klasen, M., König, A.A., Roos, R., Gaebler, A.J., Mathiak, K., 2019. Fronto-parietal and temporal brain dysfunction in depression: A fMRI investigation of auditory mismatch processing. Hum. Brain Mapp. 40, hbm.24623. https://doi.org/10.1002/hbm.24623.