Comparing the Climatic and Landscape Risk Factors for Lyme Disease Cases in the Upper Midwest and Northeast United States

Yuting Dong 1, Zheng Huang 1,* Yong Zhang 2, Yingying X.G. Wang 3 and Yang La 4,*

1 College of Life Sciences, Nanjing Normal University, Nanjing 210046, China
2 College of Biology and the Environment, Nanjing Forestry University, Nanjing 210037, China
3 Department of Biological and Environmental Science, University of Jyväskylä, FI-40014 Jyväskylä, Finland
4 Medical College, Tibet University, Lhasa 850000, China
* Correspondence: 08313@njnu.edu.cn (Z.H.); yangla721@utibet.edu.cn (Y.L.)

Received: 29 January 2020; Accepted: 23 February 2020; Published: 28 February 2020

Abstract: Lyme disease, recognized as one of the most important vector-borne diseases worldwide, has been increasing in incidence and spatial extend in United States. In the Northeast and Upper Midwest, Lyme disease is transmitted by *Ixodes scapularis*. Currently, many studies have been conducted to identify factors influencing Lyme disease risk in the Northeast, however, relatively few studies focused on the Upper Midwest. In this study, we explored and compared the climatic and landscape factors that shape the spatial patterns of human Lyme cases in these two regions, using the generalized linear mixed models. Our results showed that climatic variables generally had opposite correlations with Lyme disease risk, while landscape factors usually had similar effects in these two regions. High precipitation and low temperature were correlated with high Lyme disease risk in the Upper Midwest, while with low Lyme disease risk in the Northeast. In both regions, size and fragmentation related factors of residential area showed positive correlations with Lyme disease risk. Deciduous forests and evergreen forests had opposite effects on Lyme disease risk, but the effects were consistent between two regions. In general, this study provides new insight into understanding the differences of risk factors of human Lyme disease risk in these two regions.

Keywords: lyme disease; *Borrelia burgdorferi*; forest fragmentation; climate

1. Introduction

Lyme disease, caused by spirochete *Borrelia burgdorferi* sensu stricto (*B. burgdorferi* hereafter), is recognized as one of the most important vector-borne diseases in United States [1,2]. Since Lyme disease was first reported in Connecticut in 1975 [3,4], it has been increasing in incidence and spatial extend in United States [5,6]. Now, Lyme disease is endemic in the Northeast, Upper Midwest and West Coast [1]. In the Northeast and Upper Midwest, Lyme disease is vectored by deer ticks (*Ixodes scapularis*), which maintain *B. burgdorferi* in a horizontal transmission cycle between ticks and multiple vertebrate hosts [7]. Disease ecologists have made great efforts to understand the transmission processes of *B. burgdorferi* and identified many biotic and abiotic risk factors that attribute to Lyme disease expansion and spread in United States [1,3,8], and these efforts have yielded a wide range of control strategies. However, the number of Lyme cases have steadily increased, with about 30,000 cases of Lyme disease (according to CDC reports) occurred annually now in United States [9]. As currently no human vaccines are available [10], a better understanding of the epidemiology and risk factors of Lyme disease is still needed.

As the process of Lyme disease spread involves hosts, vectors and pathogens, any factors that can potentially influence their survivals, distributions and movements may affect the risk of disease.
transmission [11,12]. Previous studies have identified many climatic and landscape factors that may attribute to Lyme disease risk [1,13–16]. For climatic factors, laboratory studies had shown that ticks are highly vulnerable to desiccation and generally had high mortality in conditions with low humidity and high temperature [17,18]. Thus, temperature and humidity may affect Lyme disease risk indirectly through the impacts on tick survivals and population dynamics [19–21]. For example, when investigating Lyme incidence in seven northeastern states, Subac found a positive relationship between disease incidence and the June moisture index in previous years [22]. This result might be explained by a later field work study which showed that heavy precipitation in late spring or early summer precipitation was the most favorable climatic factor for tick survival in the Northeast [21]. Besides precipitation and humidity, temperature has also been correlated to Lyme disease risk. A recent study exploring the county-level Lyme spread across the United States found that the mean temperature was negatively correlated with Lyme disease spread [6], which was consistent to a previous study which also showed a negative correlation between the county-level Lyme incidence and the maximum annual temperature in the Northeast [3].

For landscape factors, a previous review had suggested that the presence of forest was consistently associated with increased Lyme disease risk [1]. Besides, forest habitat configurations can also be important in affecting Lyme disease risk due to its impacts on host movements and distributions, as well as the contact rates between human and ticks [23–26]. Human activity like urbanization induced fragmentation, increasing the amount of edge habitats between residential development and forests [3,27]. These edge habitats serve as preferred habitats for many host species of ticks, particularly the white-tailed deer that is the main host for adult ticks [28], and thus can increase the entomological risk of Lyme disease [29]. Forest fragmentation may also increase the contact rates between human population and ticks, which can elevate human exposure to Lyme disease [30]. However, there is also a different mechanism, suggesting that the spread of pathogens and tick vectors may be slowed down in fragmented patches due to the restriction on host movements [31].

When retrospecting studies on the risk factors of Lyme disease in the United States, we may find that relatively fewer studies focused on the Upper Midwest, comparing to the Northeast. It has been suggested that Lyme disease in these two regions originated from different places (Connecticut for the Northeast, and Wisconsin for the Upper Midwest) [4]. Besides, the seasonality in tick feeding also showed some differences, though B. burgdorferi is typically transmitted by the same tick species I. scapularis in these two regions. In the Northeast, nymphs feed predominantly during May and July, and larvae mainly take their bloodmeals from June to September, while the seasonal timing of larval and nymphal feeding coincide in the Upper Midwest [32]. This seasonal synchrony in nymphal and larval feeding may make the Lyme dynamics and risk factors different to those in the Northeast. In this study, we explore the climatic and landscape factors that influence the spatial patterns of Lyme cases and compare the risk factors in the Northeast and Upper Midwest United States. Our results suggested that climatic variables generally showed opposite correlations with Lyme disease risk, while landscape factors usually had similar effects in these two regions.

2. Materials and Methods

2.1. Lyme Disease Data

The study area (Figure 1) includes 13 states in the Northeast (Connecticut, Delaware, Maine, Maryland, Massachusetts, New Hampshire, New Jersey, New York, North Carolina, Pennsylvania, Rhode Island, Vermont, and Virginia; not all of these states are considered to be in the Northeast, but here we follow a previous study [3], including all 13 states due to their geographical contiguity and high Lyme incidence) and six states in the Upper Midwest (Illinois, Indiana, Iowa, Michigan, Minnesota, Wisconsin) of United States. The annual number of human Lyme disease cases for each county during 2012–2016 were obtained from the Centers for Disease Control and Prevention
According to a previous study [11], we limited our study area to those counties with established or reported *I. scapularis* populations.

**Figure 1.** Map of study area. A—the Upper Midwest; B—the Northeast of United States.

### 2.2. Data of Predictors

For each county in each year, we calculated the mean temperature (MeanTem), maximum temperature (MaxTem) and mean precipitation (Pre) of each season (spring, summer, autumn, and winter) in previous year (Table 1), based on the Climate Research Unit (CRU) datasets [33], a time-series dataset that yields month-by-month variations in climate. The processing of climatic data was carried out in ArcGIS 10.2.2.

**Table 1.** Description of climatic and landscape factors used in this study.

| Predictors | Descriptions |
|------------|--------------|
| Pre_1      | Mean precipitation in previous spring |
| Pre_2      | Mean precipitation in previous summer |
| Pre_3      | Mean precipitation in previous autumn |
| Pre_4      | Mean precipitation in previous winter |
| MeanTem_1  | Mean temperature in previous spring |
| MeanTem_2  | Mean temperature in previous summer |
| MeanTem_3  | Mean temperature in previous autumn |
| MeanTem_4  | Mean temperature in previous winter |
| MaxTem_1   | Mean maximum temperature in previous spring |
| MaxTem_2   | Mean maximum temperature in previous summer |
| MaxTem_3   | Mean maximum temperature in previous autumn |
| MaxTem_4   | Mean maximum temperature in previous winter |
| CA_X       | Total area of a land cover class X |
| PLAND_X    | Percentage of area of a land cover class X |
| TE_X       | Total edge length of a land cover X at the region |
| ED_X       | Edge density of a land cover X at the region |
| DIST_O     | Distance to the origin area of Lyme disease |

Land cover data of 2013 was accessed from the National Land Cover Database (NLCD) [34]. Following a previous study [3], we focused on seven particular land cover classes: deciduous
forest (class 41), evergreen forest (class 42), mixed forest (class 43), developed-open space (class 21), developed-low intensity space (class 22), developed-medium intensity space (class 23), and developed-high intensity space (class 24). For each county, we then derived several landscape indicators for each land cover class (Table 1), including CA (total area of a specific land cover class), PLAND (percentage of a land cover respect to the total county area), TE (total edge length), ED (edge density, total edge length divided by the total county area). Following a previous study [5], we also include, in addition to climatic and landscape predictors, the distance to the origin areas of Lyme disease (Connecticut for Northeast and Wisconsin for Upper Midwest). The processing of landscape data was carried out in ArcGIS 10.2.2 and Fragstats 4.2.

2.3. Statistical Analyses

Following previous studies [3,5,35], we applied generalized linear mixed models (GLMM) with negative binomial regression to investigate the relationships between Lyme disease cases and predictors, as negative binomial regression allows for the overdispersion that was commonly encountered in reported cases of Lyme disease [29,36]. We included state and year as random factors to control for the variations between years and states. Before performing GLMMs, we scaled all predictor variables to have a mean of zero and a standard deviation of one.

With GLMMs, we first conducted univariate regression analyses to test the association of each predictor with Lyme disease risk. Predictors with a \( p \)-value < 0.05 were identified as potential risk factors which were used to conduct model averaging. Before performing model averaging, we checked for the multicollinearity by examining the correlation coefficients (\( r \)) between potential risk factors. For highly correlated factors (\( r > 0.7 \)) [37], we only included the variable with the smaller \( p \)-value in model averaging. After removing highly correlated predictors, we constructed a full model with all remained potential risk factors. Based on the changed Akaike information criterion (AICc) values [38], we then ranked the candidate models and considered the models within \( \Delta \text{AICc} < 2 \) as competing models, which were used to average the regression coefficient of each predictor variable. For both univariate analyses and model averaging analyses, the county area (AREA) was retained in the model to control for the effect of area size. All statistical analyses were conducted in RStudio® version 1.1.463 (RStudio, Inc., Boston, MA, USA) with lme4 [39] and MuMIn [40] packages.

3. Results

3.1. Univariate Regression Analyses

Our results from univariate analyses (Table 2) showed that the distance to original disease area (Dist_O) had a negative correlation with Lyme cases in both the Northeast and Upper Midwest. The mean summer precipitation in previous year (Pre_2) was positively correlated with Lyme cases in Upper Midwest, while the mean autumn precipitation (Pre_3) was negatively correlated with Lyme disease risk in Northeast. The seasonal maximum temperature in previous year generally had a better predictive power than the seasonal mean temperature. The maximum temperature generally had negative effects on Lyme disease risk in Upper Midwest, while had positive effects in Northeast.
Table 2. Summary (Mean ± S.D.) and univariate regression results (standardized regression coefficient, b, and t) for the predictors correlated with the Lyme cases in the Northeast and Upper Midwest United States.

| Variables | Upper Midwest | | Northeast | | |
|-----------|---------------|-------|-----------|-------|
|           | Mean ± S.D.   | b     | t         | Mean ± S.D. | b     | t      |
| Dist_O    | 125 ± 157     | −1.39 | −12.9 **  | 296 ± 248   | −1.25 | −12.5 *** |
|           |               |       |           |               |       |        |
| Pre_1     | 100 ± 35.4    | −0.073| −1.0      | 90.6 ± 29.3  | −0.056| −1.15  |
| Pre_2     | 82.2 ± 24.5   | 0.23  | 3.51 ***  | 114 ± 38.7   | −0.006| −0.995 |
| Pre_3     | 72.0 ± 29.6   | −0.072| −1.02     | 87.0 ± 26.7  | −0.098| −2.09 *|
| Pre_4     | 54.9 ± 23.4   | −0.56 | −5.62 *** | 85.9 ± 18.7  | 0.067 | 1.13   |
| MeanTem_1 | 13.4 ± 5.41   | −0.066| −1.07     | 13.2 ± 5.50  | 0.024 | 0.52   |
| MeanTem_2 | 24.4 ± 3.43   | −0.15 | −2.56 *   | 24.3 ± 3.76  | 0.076 | 1.67   |
| MeanTem_3 | 13.9 ± 4.36   | −0.039| −0.66     | 13.8 ± 4.36  | 0.003 | 0.07   |
| MeanTem_4 | 0.39 ± 6.99   | 0.008 | 0.13      | 2.41 ± 6.83  | −0.016| −0.36  |
| MaxTem_1  | 15.1 ± 4.02   | −1.21 | −8.54 *** | 17.4 ± 3.88  | 0.58  | 4.12   |
| MaxTem_2  | 28.6 ± 2.07   | −0.74 | −6.47 *** | 28.6 ± 2.39  | 0.66  | 6.61   |
| MaxTem_3  | 16.40 ± 2.29  | −0.82 | −7.62 *** | 18.3 ± 3.02  | 0.44  | 3.73   |
| MaxTem_4  | 2.55 ± 4.67   | −1.33 | −9.32 *** | 6.25 ± 5.18  | 0.041 | 0.29   |

Note: * p < 0.05; ** p < 0.01; *** p < 0.001; 1 Pre_X, seasonal mean precipitation in previous year, X—(1, spring; 2, summer; 3, autumn; 4, winter); MeanTem_X, mean temperature in pervious year; MaxTem, mean maximum temperature in previous year; 2 CA_X, Total area of a land cover class X, X—(21, developed-open space; 22, developed-low intensity space; 23, developed-medium intensity space; 24, developed-high intensity space; 41, deciduous forest; 42, evergreen forest; 43, mixed forest); PLAND_X, Percentage of area of a land cover class X; TE_X, Total edge length of a land cover X; ED_X, Edge density of a land cover X.

For landscape predictors (Table 2), all four indicators (CA, PLAND, TE, ED) related to the developed area (land cover class: 41,42,43,44) generally had positive effects on Lyme disease risk in both the Upper Midwest and Northeast. The percentage of deciduous forest (PLAND_41) showed...
negative correlation with Lyme disease risk in the Upper Midwest, while the total edge length (TE_41) and edge density (ED_41) showed positive correlations in the Northeast. For evergreen forest, all four indicators had negative effects on Lyme disease risk in both regions. For mixed forest, the total forest area (CA_43), the percentage of forest area (PLAND_43) and the total edge length (TE_43) were negatively associated with Lyme disease risk in the Upper Midwest, while only the CA_43 had a significant negative effect in the Northeast.

3.2. Model Averaging Analyses

The results of model averaging (Table 3) showed that the distance to original disease area (Dist_O) had a negative correlation with Lyme cases in both the Northeast and Upper Midwest. Besides of Dist_O, the total edge length (TE_21), the edge density of open space developed area (ED_21), and the percentage of deciduous forests (PLAND_21) were positively associated with Lyme cases in the Upper Midwest. In the Northeast, the total edge length of low intensity developed area (TE_22), the edge density of deciduous forests (ED_41) and the percentage of high intensity developed area (PLAND_24) had positive effects on Lyme disease risk; while the percentage of evergreen forests (PLAND_42) have a negative effect.

Table 3. Summary statistics (averaged regression coefficient, b, Z-statistics, and p-values) for the predictors correlated with the Lyme cases in model averaging in the Northeast and Upper Midwest United States. For explanation of the variables, see Table 1.

| Variables | Upper Midwest | | | Northeast | | |
| --- | --- | --- | --- | --- | --- | --- |
| | b | Z | p-Value | b | Z | p-Value |
| Dist_O | −1.12 | 12.8 *** | <0.001 | −0.60 | 5.02 *** | <0.001 |
| PRE_2 | −0.004 | 0.22 | 0.827 | PRE_3 | −0.003 | 0.21 | 0.831 |
| MeanTem_2 | 0.02 | 0.565 | 0.827 | MaxTem_2 | 0.14 | 2.07 * | 0.038 |
| MaxTem_4 | 0.001 | 0.073 | 0.942 | TE_21 | 0.42 | 4.66 *** | <0.001 |
| TE_22 | 0.36 | 5.09 *** | <0.001 |
| PLAND_22 | 0.05 | 1.10 | 0.270 | PLAND_24 | 0.16 | 4.85 *** | <0.001 |
| PLAND_41 | 0.34 | 6.65 *** | <0.001 | TE_41 | 0.060 | 0.469 | 0.638 |
| ED_41 | 0.036 | 0.34 | 0.731 |
| CA_42 | −0.002 | 0.118 | 0.906 | PLAND_42 | 0.036 | 3.02 ** | 0.002 |
| PLAND_43 | 0.003 | 0.165 | 0.869 |

Note: * p < 0.05; ** p < 0.01, *** p < 0.001. 
1 Pre_X, seasonal mean precipitation in previous year, X—(2, summer; 3, autumn; 4, winter); MeanTem_X, mean temperature in pervious year; MaxTem, mean maximum temperature in previous year. 
2 CA_X, Total area of a land cover class X, X—(21, developed-open space; 22, developed-low intensity space; 24, developed-high intensity space; 41, deciduous forest; 42, evergreen forest; 43, mixed forest); PLAND_X, Percentage of area of a land cover class X; TE_X, Total edge length of a land cover X; ED_X, Edge density of a land cover X.

4. Discussion

In this study, we explored the correlations of climatic and landscape factors with the Lyme cases at county level in the Northeast and Upper Midwest United States. The results from univariate analyses suggested that the landscape factors related to developed area and forests generally had similar effects on Lyme disease risk in the two regions. In contrast, climatic factors generally showed opposite
relationships with Lyme disease risk in the two regions. The results from model averaging analyses in two regions only identified several but quite different risk factors. As many climatic and landscape factors were highly correlated with each other, the significant effect of a specific factor in multiple models might also be caused by other highly correlated factors. Therefore, we here focus more on discussing the results from univariate analyses.

In both regions, the seasonal mean maximum temperature in previous year were better than the mean temperature in previous year in explaining the spatial patterns of Lyme cases. Increasing the mean maximum temperature in previous year was associated with a decrease in the number of Lyme cases in Upper Midwest, while associated with an increase in Lyme disease risk in Northeast. The precipitation in previous summer was positively correlated with Lyme disease risk in Upper Midwest, while the precipitation in previous autumn showed a negative association in Northeast. The results from the Upper Midwest seems consistent to the expectation that low humidity and high temperatures could regulate tick abundance [21,22]. In contrast, the results from the Northeast conflicted with this expectation, but consistent with a previous study which also suggested a tick abundance when there was a high temperature at ground level [41]. These results confirmed the conclusion from a previous study which suggest that the effects of weather variables can vary considerably among different regions [42].

In contrast to climatic factors, most landscape factors showed similar effects on Lyme disease risk in the Northeast and Upper Midwest. Both the area size related indices (CA and PLAND) and fragmentation indices (TE and ED) of developed area (land cover class: 21–24) showed very strong positive correlations with Lyme disease risk (as seen in Table 2). As these indices were generally positively correlated with each other, we could not draw the conclusion which factors had true causal effects on Lyme disease risk. However, we found that in both regions, the multiple regression models included the fragmentation related indices of developed area (TE22 for Northeast; TE21 and ED21 for Upper Midwest; see Table 3), which might indicate that Lyme disease risk generally increased in fragmented developed area. These results were consistent to a previous study [3]. According to the NLCD 2013 classification, the open developed area (land cover class 21) and the low intensity developed area (class 22) are most likely single family housing units. The fragmentation of these types of land covers indicated a high chance of the occurrence of surrounding forests or herbaceous cover. Therefore, the contact rates between human and ticks might be enhanced in these areas [3]. Besides, edge habitats of residential area usually can provide more food resources for white tailed deer, the major host for adult ticks, increasing tick abundance [29]. Both of these two mechanisms could result in a higher Lyme disease risk in fragmented residential habitats.

The fragmentation of deciduous forests generally increases the number of Lyme cases (see Table 2). Previous studies have proposed that tick abundance is generally higher in fragmented deciduous forests, as forest fragmentations may provide ideal habitats for many reservoir hosts of ticks [1,30]. In fact, it had been shown that the entomological risk of Lyme disease risk was usually higher in small forest fragments due to the high abundance of white-footed mouse [29,30,43]. Moreover, edges in fragmented forests might be utilized more frequently by humans, resulting in higher contact rates between human and infected ticks [23]. After controlling for other factors in multiple regression models, the percentage of area of deciduous forests (PLAND 41) also had a positive effect in the Upper Midwest, consistent with many previous studies that had demonstrated the important role of forest cover in determining Lyme disease risk at landscape level. These studies suggested that more forest generally means more habitats for hosts, providing the blood meals for ticks, and thus the density of infected questing ticks [8,12,28]. In contrast to deciduous forests, the number of Lyme cases was lower in evergreen forests (class 42) in both regions. These results were also consistent to a previous study that suggested evergreen forests were located in mountainous areas, poor environments for ticks regarding to temperature and precipitation [3].

We must admit that the Lyme case number obtained from CDC might be an underestimate of actual human cases [1]. Particularly, different states may apply different approaches to gather case
data. Including state as a random effect in our analyses was able to control for, to some extent, the differences in surveillance way among states.

5. Conclusions

In this study, we explored and compared the roles of climatic and landscape factors in shaping the spatial patterns of Lyme cases in the Upper Midwest and Northeast United States. Our results showed that climatic variables generally showed opposite correlations with Lyme disease risk, while landscape factors usually had similar effects in these two regions. High Lyme disease risk was correlated with high precipitation and low temperature in the Upper Midwest, while with low precipitation and high temperature in the Northeast. In both regions, area size related and fragmentation related indices of developed area showed strong positive correlations with Lyme disease risk. Deciduous forests and evergreen forests had opposite effects on Lyme disease risk, but the effects were consistent between two regions. Our study is the first study, to our knowledge, comparing the risk factors for Lyme disease in the Upper Midwest and the Northeast, and thus may provide new insight into understanding the differences of risk factors of Lyme disease risk in these two regions.

Author Contributions: Data collection, Y.D., Y.X.G.W. and Y.L.; Data processing, Y.D. and Z.H.; Data analyses, Y.D., Y.Z. and Y.X.G.W.; Supervision, Z.H. and Y.L.; Writing—original draft, Y.D. and Y.L.; Writing—review & editing, Z.H., Y.Z. and Y.X.G.W. All authors have read and agreed to the published version of the manuscript.

Funding: This research was funded by National Natural Science Foundation of China, grant number 31870400; and the Cultivation project of Tibet University.

Conflicts of Interest: The authors declare no conflict of interest. The funders had no role in the design of the study; in the collection, analyses, or interpretation of data; in the writing of the manuscript, or in the decision to publish the results.

References

1. Killilea, M.E.; Swei, A.; Lane, R.S.; Briggs, C.J.; Ostfeld, R.S. Spatial dynamics of Lyme disease: A review. *EcoHealth* **2008**, *5*, 167–195. [CrossRef] [PubMed]
2. McClure, M.; Diuk-Wasser, M. Reconciling the entomological hazard and disease risk in the Lyme disease system. *Int. J. Environ. Res. Public Health* **2018**, *15*, 1048. [CrossRef] [PubMed]
3. Tran, P.M.; Waller, L. Effects of landscape fragmentation and climate on Lyme disease incidence in the northeastern United States. *EcoHealth* **2013**, *10*, 394–404. [CrossRef] [PubMed]
4. Steere, A.C.; Coburn, J.; Glickstein, L. The emergence of Lyme disease. *J. Clin. Investig.* **2004**, *113*, 1093–1101. [CrossRef]
5. Turney, S.; Gonzalez, A.; Millien, V. The negative relationship between mammal host diversity and Lyme disease incidence strengthens through time. *Ecology* **2014**, *95*, 3244–3250. [CrossRef]
6. Wang, Y.X.; Matson, K.D.; Xu, Y.; Prins, H.H.; Huang, Z.Y.; de Boer, W.F. Forest connectivity, host assemblage characteristics of local and neighboring counties, and temperature jointly shape the spatial expansion of Lyme disease in United States. *Remote Sens.* **2019**, *11*, 2354. [CrossRef]
7. Barbour, A.G.; Fish, D. The biological and social phenomenon of Lyme disease. *Science* **1993**, *260*, 1610–1616. [CrossRef]
8. Wood, C.L.; Lafferty, K.D. Biodiversity and disease: A synthesis of ecological perspectives on Lyme disease transmission. *Trends Ecol. Evol.* **2013**, *28*, 239–247. [CrossRef]
9. DeLong, A.; Hsu, M.; Kotsoris, H. Estimation of cumulative number of post-treatment Lyme disease cases in the US, 2016 and 2020. *BMC Public Health* **2019**, *19*, 352. [CrossRef]
10. Sanchez, E.; Vannier, E.; Wormser, G.P.; Hu, L.T. Diagnosis, treatment, and prevention of Lyme disease, human granulocytic anaplasmosis, and babesiosis: A review. *JAMA* **2016**, *315*, 1767–1777. [CrossRef]
11. Estrada-Peña, A.; Ostfeld, R.S.; Peterson, A.T.; Poulin, R.; de la Fuente, J. Effects of environmental change on zoonotic disease risk: An ecological primer. *Trends Parasitol.* **2014**, *30*, 205–214. [CrossRef]
12. Huang, Z.Y.X.; van Langevelde, F.; Prins, H.H.T.; de Boer, W.F. The diversity–disease relationship: Evidence for and criticisms of the dilution effect. *Parasitology* **2016**, *143*, 1075–1086. [CrossRef]
13. Bron, G.M.; del Pilar Fernandez, M.; Larson, S.; Maus, A.; Gustafson, D.; Tsao, J.I.; Diuk-Wasser, M.A.; Bartholomay, L.C.; Paskewitz, S.M. Context matters: Contrasting behavioral and residential risk factors for Lyme disease between two high-incidence regions in the Northeastern and Midwestern US. medRxiv 2020. [CrossRef]
14. Sharareh, N.; Behler, R.P.; Roome, A.B.; Shepherd, J.; Garruto, R.M.; Sabounchi, N.S. Risk Factors of Lyme Disease: An Intersection of Environmental Ecology and Systems Science. Healthcare 2019, 7, 66. [CrossRef] [PubMed]
15. Moon, K.A.; Pollak, J.; Poulsen, M.N.; Hirsch, A.G.; DeWalle, J.; Heaney, C.D.; Aucott, J.N.; Schwartz, B.S. Peridomestic and community-wide landscape risk factors for Lyme disease across a range of community contexts in Pennsylvania. Environ. Res. 2019, 178, 108649. [CrossRef] [PubMed]
16. Fischhoff, I.R.; Bowden, S.E.; Keesing, F.; Ostfeld, R.S. Systematic review and meta-analysis of tick-borne disease risk factors in residential yards, neighborhoods, and beyond. BMC Infect. Dis. 2019, 19, 1–11.
17. Needham, G.R.; Teel, P.D. Off-Host Physiological Ecology of Ixodid Ticks. Annu. Rev. Entomol. 1991, 36, 659–681. [CrossRef] [PubMed]
18. McCabe, G.J.; Burnell, J.E. Precipitation and the occurrence of Lyme disease in the northeastern United States. Vector. Borne. Zoonotic. Dis. 2004, 4, 143–148. [CrossRef] [PubMed]
19. Subak, S. Effects of climate on variability in Lyme disease incidence in the northeastern United States. Am. J. Epidemiol. 2003, 157, 531–538. [CrossRef] [PubMed]
20. Horobik, V.; Keesing, F.; Ostfeld, R.S. Abundance and Borrelia burgdorferi-infection prevalence of nymphal Ixodes scapularis ticks along forest–field edges. EcoHealth 2006, 3, 262–268. [CrossRef]
21. Millins, C.; Dickinson, E.R.; Isakovic, P.; Gilbert, L.; Wojciechowska, A.; Paterson, V.; Tao, F.; Jahn, M.; Killbride, E.; Birtles, R. Landscape structure affects the prevalence and distribution of a tick-borne zoonotic pathogen. Parasite Vector. 2018, 11, 1–11. [CrossRef] [PubMed]
22. VanAcker, M.C.; Little, E.A.; Molaei, G.; Bajwa, W.I.; Diuk-Wasser, M.A. Enhancement of Risk for Lyme Disease by Landscape Connectivity, New York, New York, USA. Emerg. Infect. Dis. 2019, 25, 1136. [CrossRef] [PubMed]
23. Sharareh, N.; Sabounchi, N.S.; Roome, A.; Spalthis, R.; Garruto, R.M. Model-based risk assessment and public health analysis to prevent Lyme disease. R. Soc. Open Sci. 2017, 4, 170841. [CrossRef]
24. Súzan, G.; Esponda, F.; Carrasso-Hernández, R.; Aguirre, A. Habitat fragmentation and infectious disease ecology. In New Directions in Conservation Medicine: Applied Cases of Ecological Health; Aguirre, A., Ostfeld, R., Daszak, P., Eds.; Oxford University Press: New York, NY, USA, 2012; pp. 135–150.
25. Ferrell, A.M.; Brinkerhoff, R.J. Using landscape analysis to test hypotheses about drivers of tick abundance and infection prevalence with Borrelia burgdorferi. Int. J. Environ. Res. Public Health 2018, 15, 737. [CrossRef]
26. Brownstein, J.S.; Skelly, D.K.; Holford, T.R.; Fish, D. Forest fragmentation predicts local scale heterogeneity of Lyme disease risk. Oecologia 2005, 146, 469–475. [CrossRef]
27. Frank, D.H.; Fish, D.; Moy, F.H. Landscape features associated with Lyme disease risk in a suburban residential environment. Landscape Ecol. 1998, 13, 27–36. [CrossRef]
28. Li, S.; Hartemink, N.; Speybroeck, N.; Vanwanmeke, S.O. Consequences of landscape fragmentation on Lyme disease risk: A cellular automata approach. PloS ONE 2012, 7, e96912. [CrossRef]
29. Eisen, R.J.; Piesman, J.; Zielinski-Gutierrez, E.; Eisen, L. What do we need to know about disease ecology to prevent Lyme disease in the northeastern United States? J. Med. Entomol. 2012, 49, 11–22. [CrossRef] [PubMed]
30. Harris, I.; Jones, P.D.; Osborn, T.J.; Lister, D.H. Updated high-resolution grids of monthly climatic observations—the CRU TS3. 10 Dataset. Int. J. Climatol. 2014, 34, 623–642. [CrossRef] [PubMed]
31. Wickham, J.; Homer, C.; Vogelmann, J.; McKerrow, A.; Mueller, R.; Herold, N.; Coulston, J. The multi-resolution land characteristics (MRLC) consortium—20 years of development and integration of USA national land cover data. Remote Sens. 2014, 6, 7424–7441. [CrossRef]
35. Bown, K.J.; Begon, M.; Bennett, M.; Woldehiwet, Z.; Ogden, N.H. Seasonal dynamics of Anaplasma phagocytophila in a rodent-tick (Ixodes trianguliceps) system, United Kingdom. Emerg. Infect. Dis. 2003, 9, 63. [CrossRef] [PubMed]
36. Jackson, L.E.; Hilborn, E.D.; Thomas, J.C. Towards landscape design guidelines for reducing Lyme disease risk. Int. J. Epidemiol. 2006, 35, 315–322. [CrossRef]
37. Zuur, A.F.; Ieno, E.N.; Elphick, C.S. A protocol for data exploration to avoid common statistical problems. Methods Ecol. Evol. 2010, 1, 3–14. [CrossRef]
38. Burnham, K.P.; Anderson, D.R.; Huyvaert, K.P. AIC model selection and multimodel inference in behavioral ecology: Some background, observations, and comparisons. Behav. Ecol. Sociobiol. 2011, 65, 23–35. [CrossRef]
39. Bates, D.; Sarkar, D.; Bates, M.D.; Matrix, L. The lme4 package. R. Package Version 2007, 2, 74.
40. Barton, K.; Barton, M.K. Package ‘MuMIn’. 17 November 2019. Available online: https://cran.r-project.org/web/packages/MuMIn/MuMIn.pdf (accessed on 17 November 2019).
41. Werden, L.; Barker, I.K.; Bowman, J.; Gonzales, E.K.; Leighton, P.A.; Lindsay, L.R.; Jardine, C.M. Geography, deer, and host biodiversity shape the pattern of Lyme disease emergence in the Thousand Islands archipelago of Ontario, Canada. PLoS ONE 2014, 9, e85640. [CrossRef]
42. Schauber, E.M.; Ostfeld, R.S.; Evans, J.; Andrew, S. What is the best predictor of annual Lyme disease incidence: Weather, mice, or acorns? Ecol. Appl. 2005, 15, 575–586. [CrossRef]
43. Allan, B.F.; Keesing, F.; Ostfeld, R.S. Effect of forest fragmentation on Lyme disease risk. Conserv. Biol. 2003, 17, 267–272. [CrossRef]