Coordinated Transcription and Decay Messenger Half-Lives and Reveals Sequential Waves of Co-ordinated Transcription and Decay

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Abstract
The transcriptome in a cell is finely regulated by a large number of molecular mechanisms able to control the balance between mRNA production and degradation. Recent experimental findings have evidenced that fine and specific regulation of degradation is needed for proper orchestration of a global cell response to environmental conditions. We developed a computational technique based on stochastic modeling, to infer condition-specific individual mRNA half-lives directly from gene expression time-courses. Predictions from our method were validated by experimentally measured mRNA decay rates during the intraerythrocytic developmental cycle of Plasmodium falciparum. We then applied our methodology to publicly available data on the reproductive and metabolic cycle of budding yeast. Strikingly, our analysis revealed, in all cases, the presence of periodic changes in decay rates of sequentially induced genes and co-ordination strategies between transcription and degradation, thus suggesting a general principle for the proper coordination of transcription and degradation machinery in response to internal and/or external stimuli.

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Introduction
Appropriate and timely changes in gene expression are essential for cell life. The transcriptome is finely regulated by a large number of molecular mechanisms able to adjust the balance between mRNA production and degradation. Every aspect of transcript life is subject to elaborate control but, traditionally, the focus of the research has been on transcriptional regulation [1]. However, whereas mRNA abundance results from the dynamic interplay between transcription and degradation, the speed by which cells can adjust their mRNA levels is critically dependent on the rate of mRNA turnover [2]. As a result, small changes in mRNA stability may dramatically drive rapid variations of transcript abundance. Efforts to understand the underlying principles of mRNA decay and transcription co-ordination are very important since the balance between transcription and decay influences most, if not all, the cell responses to endogenous and exogenous signals [3].

The current widespread interest in this topic has been fostered by the finding of specific regulatory mechanisms of mRNA stability such as, for example, RNA binding proteins [4,5] and small RNAs [6]. Regulation of transcript stability cannot be considered a simple “disposal system” but a sophisticated tool for the proper orchestration of the global cell response to internal and external stimuli [6]. Remarkably, a key role of mRNA stability has been reported in cancer, inflammatory diseases and Alzheimer's [7]. In recent years there has been a surge in empirical studies that measured, on a genome-wide scale in a variety of environmental conditions, messenger half-lives of many organisms, including plants [8] mammals [9] and fungi [2]. The discovery of such new regulatory layer has clarified that, in order to obtain a clear picture of the underlying regulatory machinery, it is necessary to complement the traditional time-course experiment measuring the cell transcriptional response under certain conditions far from steady state with decay rates data under the same condition [10].

Experimental procedures for the evaluation of mRNA decay rates are based on measuring gene expression upon inhibition of transcription [11–13] or on pulse-chase RNA labeling protocols [2,13–15]. Such protocols are very critical (see Figure 1 for a comparison among different studies summarized in Table 1), since, for instance, transcriptional shut-off blocks growth and has a profound effect on cellular physiology, as well as on mRNA metabolism [2]. In fact, Wang et al. [11] and Grigull et al. [12] datasets show a low value of the Pearson correlation ($r \approx 0.36$), and no correlation at all can be found ($r \approx 0.01$) between Munchel et al. [2] and Wang et al. [11] datasets (see Figure 1A and Figure 1B respectively). Despite the same experimental conditions (asynchronous growth), the two half-life independent measurements obtained
expression kinetics to identify condition-specific mRNA stabilities (datasets; correlation between Munchel increase in correlation between Wang Consistently, by averaging using functional groups, we found an function [11,12] are likely to share similar half-life values.

by Wang et al. and Munchel et al. are uncorrelated, probably due to differences in the shut off protocol (pulse chase for [2] and thermal inactivation for [11]), whereas Grigull et al. and Wang et al. appears significantly correlated, probably due to the same shut off protocol used (thermal inactivation).

It has been shown that genes having the same biological function [11,12] are likely to share similar half-life values. Consistently, by averaging using functional groups, we found an increase in correlation between Wang et al. and Grigull et al. datasets ($r=0.64$, see Supplementary Figure S1A), and still no correlation between Munchel et al. and Wang et al. datasets ($r=0.0043$, see Supplementary Figure S1B).

Here, we developed a stochastic computational model of the expression kinetics to identify condition-specific mRNA stabilities which makes use only of experimental mRNA time profiles. We also assumed that degradation rates are gene-specific but approximately constant over the experiment time course. Predictions of our algorithm, termed DRAGON (Decay RAtes from Gene expression), were validated on experimental mRNA abundance [16] and turnover [17] data, both collected during the Intraerythrocytic Developmental Cycle (IDC) of Plasmodium falciparum. The estimations were in line with the experimental measurements. Remarkably, the DRAGON estimated half-lives were consistent with the finding of a peculiar pattern of mean half-life values along the wave of sequentially induced genes in subsequent stages of P. falciparum development. We also applied our methodology to public time-series datasets for which half-lives data, under the same experimental conditions, have not been experimentally measured. In particular we focused on budding yeast reproductive [18,19] and metabolic cycle data [20]. In fact, for the yeast Saccharomyces cerevisiae, only half-life data under asynchronous growth are publicly available [2,11,12]. Our study showed the presence of the same periodic pattern of mean half-life values in all datasets, thus suggesting that such behavior may be a general feature, not limited to the Plasmodium falciparum IDC.

mRNA kinetics and half-life

Experimental evidence suggest that the majority of mRNAs are degraded with a first-order decay rate [2,21]. This allows to characterize mRNA disappearance time profiles by a first-order rate equation

$$\frac{dx(t)}{dt} = P(t) - kx(t)$$

where $k$ is the decay rate (or half-life $t_{1/2}$, with $t_{1/2} = \ln(2)/k$), $x(t)$ is the mRNA concentration and $P(t)$ is the promoter activity (the rate of production of new mRNAs). It is worth noting that, the degradation rate $k$ cannot be estimated from the concentration time profile $x(t)$ for a single gene, since the term $P(t)$ is not usually available in the typical time-course microarray experiment. The measurement of the promoter activity time profile would require additional experiments (such as those described in [15]) but, in this paper, we will assume that only mRNA abundance time-series

![Figure 1. Basic comparison statistics among the yeast S. cerevisiae mRNA half-lives during asynchronous growth measured by three independent laboratories. Three genome-wide studies considered are: Grigull et al. [12], Wang et al. [11] and Munchel et al. [2]. (A) Scatterplot of Wang et al. and Grigull et al. datasets; (B) Scatterplot of Munchel et al. and Wang et al. datasets.](https://doi.org/10.1371/journal.pcbi.1002772.g001)
data are available. At steady-state $x = x_{ss}$ and $P(t) = P$ are constant so that $\frac{dx_{ss}(t)}{dt} = 0$ and, consequently, we obtain

$$x_{ss} = \frac{P}{k}$$

(2)

From the above equation, it is clear that at steady-state an increase (decrease) in mRNA concentration can be produced either by an increase (decrease) of transcription or by a decreased (increased) value of the decay rate: the two regulatory strategies have therefore an equivalent outcome. As a result, from steady-states measurements, it is hopeless to reveal the relative contribution of transcription and degradation and, most importantly, their coordinated activity as well. By contrast, the whole kinetics of induction and relaxation, as measured by time-courses experiments, depends on the degradation and production rate in different ways: increasing (decreasing) the production rate results in a proportionally increased (decreased) mRNA abundance, whereas the rise time (i.e. the time required for the response to rise from 10% to 90% of its final value) is not affected. Increasing the decay rate results in a faster rise time both in the induction and relaxation phases, whereas a decrease results in slower rise time [10]. This key point is illustrated in Figure 2A and in Supplementary Figures S2A and S2B.

Another important consequence of half-life specificity is the regulation of the timing of gene induction, as pointed out by Elkon et al. [22]. In fact, an expression wave, i.e. the sequential activation of genes, is usually interpreted as resulting from the corresponding activation of a multi-step transcription factors cascade (as illustrated in Figure 2B). Whereas such mechanism is certainly very important, there is also an alternative way to obtain an expression wave by means of a “stability gradient”. As illustrated in Figure 2C, a single transcription factor may initiate transcription of a set of target genes and their peak of induction can be modulated by a stability “gradient”, i.e. by specifically adjusted decay rates. More precisely, early induced genes would have short half-lives and late responding genes would have long half-lives. Clearly, both mechanism may well act in cells, thus generating a wide spectrum of responses.

Time-courses are a very common design for microarray analysis, which allows researchers to follow the dynamics of the cellular response to perturbations [22]. Such data are available for a very large number of experimental conditions and organisms:

![Figure 2. Kinetics of gene induction.](image)

Panel A shows in silico experiments to illustrate some basic features of gene induction kinetics. The “ON” and “OFF” regions correspond to the turning “ON” or “OFF” of the promoter activity. (A) Induction kinetic of transcripts having the same steady-state concentration, but different half-lives and synthesis rates reaching the same steady state value. The time profile plotted in red corresponds to an unstable transcript and displays a fast induction and relaxation profile. By contrast, the blue one has an higher half-life value, resulting in a slower response. (B) Cascade of transcription factors resulting in waves of sequentially induced genes. The timing of expression peaks is modulated by transcriptional serial regulation. (C) Sequentially induced genes generated by a single transcription factor and a stability “gradient”. The timing of expression peaks is modulated by post-transcriptional regulation. Early induced genes are those with a low half-life value, late induced ones are those with an high half-life value.

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only the Stanford Microarray Database includes to date 1545 time course data sets. Among the examples later illustrated in the paper, it is worth mentioning the genome-wide gene expression time-series obtained during the reproductive cell cycle [18,19], the metabolic cycle [20] and the *P. falciparum* IDC [16]. The time-series datasets used in this paper are summarized in Table 2.

DRAGON–an algorithm for half-life estimation

The goal of the DRAGON methodology is to derive a robust estimate of each mRNA species half-life starting from all available gene expression pairs. The rationale for the algorithm mainly draws on properties of pairs exhibiting a certain degree of common promoter activity (as in [23]). Besides, DRAGON infers common promoter activity using a statistical model that simulates both gene-specific and common effects.

The rate of change of mRNA concentration for a generic pair of genes, say gene *j* and gene *h*, is:

\[
\frac{dx_j(t)}{dt} = P_j(t) - k_j x_j(t) \\
\frac{dx_h(t)}{dt} = P_h(t) - k_h x_h(t)
\]  

(3)

where the symbols *x*_j*(t)* and *x*_h*(t)* represent the mRNA time profiles of the gene pair *j* and *h*, *P*_j*(t)* and *P*_h*(t)* are the promoters activity, and *k*_j* and *k*_h* are the degradation rate of mRNA of gene *j* and *h*, respectively. The terms *P*_j*(t)*, *P*_h*(t)* are not known since we considered the case in which only mRNA abundance is measured.

We modeled promoter activities as the sum of two terms, the first one common to the pair and the other one specific for each gene:

\[
P_j(t) = b_j u(t) + w_j(t) \\
P_h(t) = b_h u(t) + w_h(t)
\]  

(4)

where *u*(t) is the common part, scaled by constants *b_j* and *b_h*, whereas *w*_j*(t)* and *w*_h*(t)* are gene specific independent stochastic processes with zero mean, that is *w*_j*(t)* = *σ_j* *ν*_j*(t)*, *w*_h*(t)* = *σ_h* *ν*_h*(t)*.

Equations (3)–(4) encompass the case of:

i) *P_j*(t), *P_h*(t) fully correlated (correlation *r* = 1) for which *b_j*,*b_h* > 0 and *σ_j* = *σ_h* = 0

ii) *P_j*(t), *P_h*(t) partially correlated (correlation *r* < 1) for which *b_j*,*b_h* > 0 and *σ_j*,*σ_h* > 0

iii) *P_j*(t), *P_h*(t) un-correlated (correlation *r* = 0) for which either *b_j* = 0 or *b_h* = 0.

Equations (3)–(4) can therefore be written for all available gene pairs; thus, for a set of *N* genes, we have *N*² pairs to analyze. For each gene pair DRAGON provides an estimate of the time profile of *u*(t), of all the parameters, *k_j*, *k_h*, *b_j*, *b_h*, and the covariance matrix of the stochastic processes. For each gene we therefore have *N* estimates of the decay rate *k*, one for each pair containing that gene.

Notice that equations (3)–(4) yield a couple of linear stochastic differential equations. Since measurements of mRNA concentrations are available only at given time points, it is necessary to transform (3)–(4) in a couple of discrete stochastic equations. The exact discretization of (3)–(4) is possible since they are linear [24]. The *Kalman filter* [25] is used on the resulting discrete equations and a maximum likelihood algorithm is exploited to generate the best possible estimate of the parameters.

A complete description of the mathematical model and of the discretization and parameters estimation procedure is given in the paragraph *Stochastic modeling of expression kinetics and Kalman filtering of the Materials and Methods* section.

Results

Performance evaluation on malaria IDC experimental data

The IDC is characterized by four morphologic stages: ring, trophozoite, schizont and late schizont. The cycle begins with the red cells invasion by merozoites followed by a remodeling of the host cell in the ring stage [16]. The merozoites then develop into trophozoites. During the schizont stage, after a period of growth, the trophozoite undergoes an asexual dividing process and the parasite is ready for the next round of invasion by new merozoites (late schizont phase).

Bozdech *et al*. [16], using microarrays, measured genome-wide mRNA abundance profiles across 48 h during one cycle of *P. falciparum* IDC, collecting one sample per hour. Later on, Shock *et al*. [17] measured mRNA half-lives of 2774 transcripts of the IDC using chemical inhibitors to reach transcriptional shut-off.

The simultaneous availability of gene expression and decay data during the same biological process (IDC) represents a natural test bed for the validation of the DRAGON algorithm. Therefore, we applied DRAGON on Bozdech *et al*. dataset to obtain mRNA stability estimations (provided in Supplementary Table S6) to be compared with Shock *et al*. measurements for performance

| Table 2. Gene expression time-series experimental measurements. |

| Experiment               | Description                               | Publication        |
|--------------------------|-------------------------------------------|--------------------|
| Plasmodium IDC           | interval, synchronized by sorbitol treatments | Bozdech *et al*. (2003) |
| Yeast cell cycle         | interval, alpha-factor synchronization    | Pramila *et al*. (2006) |
| Yeast cell cycle         | interval, synchronized by centrifugal elutriation | Orlando *et al*. (2008) |
| Yeast metabolic-cycle    | interval, spontaneous synchronization      | Tu *et al*. (2005)  |
evaluation. The resulting Pearson correlation between in vitro and in silico measures is $r = 0.6$ (P value $2.4 \times 10^{-61}$), and the first principal component accounts for 82% of the variability, thus showing a good performance for DRAGON algorithm (see Figure 3). However, since gene expression and decay data have been measured by different groups, we can speculate that 18% of unexplained variability may be partly due to inherent biological variability and to transcriptional inhibition stress. As further analysis, we computed average mRNA half-lives in both studies for functional categories (see Supplementary Figure S3). We found that the two studies are in better accordance when half-lives are averaged for all genes within any given functional category (Pearson correlation $r = 0.74$).

Remarkably, Shock et al. in [17], found progressive stage-dependent average increases in mRNA stability and suggested such phenomenon to be a major determinant of mRNA accumulation (see Figure 4A). The same feature is also found using DRAGON estimated half-lives (see Figure 4B). To investigate in further detail the behavior of average half-life of genes sequentally induced during IDC, we computed for each gene the time point corresponding to its peak of expression (see the Data processing paragraph of the Materials and Methods section for details) and selected 48 groups of genes having peak of expression at each hourly time points over the 48 hours monitored by Shock et al. For each gene group we computed half-lives mean and standard deviation and found a high correlation with the corresponding curve obtained using experimental data (Pearson correlation $r = 0.8$, P value $2.1 \times 10^{-30}$; see Figure 4C). Early responding genes are characterized by high instability, whereas late responders are more stable, as also reported by Elkon et al. in [22] when studying mammalian cells. A possible explanation for the presence of stable mRNAs at the schizont stage, suggested by Shock et al., is that it may be important for the merozoite to receive a carefully regulated “starting package”, that would allow rapid activation of the IDC following the next round of invasion [17]. By contrast, the initial low mRNA stability values may be an indication of the fast dynamic remodeling after merozoite invasion [17]. To evaluate the probability of obtaining such behavior by chance, we randomized the gene expression matrix and used DRAGON to estimate half-lives (see Figure 4D). Consistently, the estimation of half-lives using random data does not produce any correlation with experimental data (Pearson correlation $r = -0.17$).

Half-lives estimation during reproductive cycle in S. cerevisiae

Gene expression during yeast cell cycle has been recently measured by Pramila et al. [18] using alpha-factor synchronization and by Orlando et al. [19] using centrifugal elutriation for synchronization. We obtained a high consistency of DRAGON estimations using data for 569 transcripts over replicate datasets (Pearson correlation $r = 0.83$ for Pramila et al. dataset and Pearson correlation $r = 0.92$ for the Orlando et al. dataset; see Figure 5A–B). The larger variability in half-lives estimations may be explained by the inconsistencies between replicate time-series in the Pramila et al. dataset with respect to the Orlando et al. dataset (see Figure 5C). All half-lives estimations obtained with the DRAGON algorithm are provided in Supplementary Tables S1 and S2 (Pramila) and in Supplementary Tables S3 and S4 (Orlando). Notwithstanding significant differences in synchronization procedures, we also found a high correlation of DRAGON half-lives estimations over the two datasets (Pearson correlation $r = 0.57$, P value $2.1 \times 10^{-38}$; see Figure 5D) where the first principal component accounts for 79% of the overall variability. We can speculate that 21% of unexplained variability may be partly due to the different synchronization methods used. In fact, Orlando et al. obtained a cell cycle duration of about 2 hours, 8 samples per cycle [19], whereas Pramila et al. obtained a cell cycle duration of about 1 hour, 12 samples per cycle. Consistently, most of the transcripts during the slower cycle display higher half-lives when compared to the fastest cycle (see Figure 5D).

GO annotations of genes with extreme half-lives in S. cerevisiae

In this paragraph we briefly discuss functional annotations (done using GOrilla software [26]) of novel predicted half-lives provided by DRAGON algorithm using yeast reproductive and metabolic cycle time series. For the yeast cell cycle we normalized the half-life log-distribution (Z-score), for each dataset, and then computed the geometric mean to obtain a single half-life value for each gene. Notably, the averaging has the effect of reducing the impact of the different synchronization stress response. The list of half-lives normalized values (geometric mean value equal to 1) for common genes to all datasets is provided as Supplementary Table S7 in the Half-life estimation paragraph of the Materials and Methods section.

Unstable genes are enriched with replication fork complex (p-value $5.95 \times 10^{-6}$) and stable genes (histones HA1-2,HB1-2) are enriched with nucleosome (p-value $7.4 \times 10^{-4}$). This is consistent with the need of producing a large number of histones during DNA replication process so that stable histone mRNAs contribute to a higher translation efficiency. Moreover, DNA replication timing requires first the formation of the replication fork, then the production of the needed histones for chromatin assembling: such temporal sequence of events is consistent with a rapid turnover of the replication complex genes and a slow turnover of the histone genes (see Supplementary Figure S4). Among unstable genes we also found the G1/S transition cyclins and among stable ones we found G2/M transition cyclins (see Supplementary Figure S5). In this case, the temporal sequence of events is the progression of the cell cycle from DNA replication to mitosis.
For the yeast metabolic cycle (half-lives estimations using DRAGON algorithm are provided in Supplementary Table S5) we found many stable mRNA species involved in the organic acid and arginine metabolism and protein catabolic processes. Among unstable messengers we found genes involved in DNA repair (p-value $4 \times 10^{-8}$), DNA metabolism (p-value $7.5 \times 10^{-8}$) and chromatin silencing (p-value $4.29 \times 10^{-3}$).

**Discussion**

**Periodic behavior of average half-lives of sequentially induced genes**

The increasing pattern of average half-life found during *P. falciparum* IDC (shown in Figure 4A) motivated us to investigate whether a periodicity could be found also in other cyclical biological processes.

We focused on the reproductive cell cycle and the metabolic cycle in *Saccharomyces cerevisiae*, for which high resolution time series measurements are available on public repositories (see Table 2).

To study if a periodic pattern of average half-life of sequentially induced genes exists along the cell cycle progression, for each gene we computed the time points at which maximal expression is attained (see the Data processing paragraph of the Material and Methods section for details). Thus, we obtained, for each time point, the list of genes having expression peak value at that time and computed the corresponding mean and variance of DRAGON estimated half-lives values. Indeed, we found a cyclic behavior along sequentially induced genes in both datasets (see Figure 6A for the Pramila et al. dataset and Figure 6B for the Orlando et al. dataset). Synchronization methods, cell cycle duration and number of samples are different between the two cited studies, but, reassuringly, the phases of the cell cycle at which mean half-life is minimal or maximal is consistent. In fact, for both datasets we observed a cyclical increase of mean half-life from G1 phase to M phase and a subsequent decrease back to G1. The figure clarifies that the minimal mean half-life is reached at the G1/S transition, whereas the maximal value correspond to the M/G2 phase for both cycles and datasets. The latter is consistent with the observation that, in higher eukaryotes, mitosis is accompanied by global repression of nuclear RNA synthesis [27], indicating that mRNAs must be stable to be inherited from daughter cells.

The yeast metabolic cycle has been recently studied by Tu et al. [20] using a continuous culture system, after a brief starvation period, the culture spontaneously began persistent respiratory cycles of about 5 hours. In the same study, a genome-wide microarray gene
expression measurement was performed. Samples were taken every 25 minutes for 3 consecutive cycles. Using DRAGON algorithm we estimated half-lives using data of 1043 transcripts. Surprisingly, also in this case we found a cyclical pattern for mean half-life of sequentially induced genes. The maximum peak is located at the RC phase and the minimum peak located at RB phase (see Figure 6C).

**Integrated analysis—sequential waves of co-ordinated transcription and decay**

Recently, the appearance of a number of studies has revealed the fundamental role of stability regulation in shaping appropriate cell response [1]. A key point has been recently addressed by Shalem et al. [10], who have shown the dynamic co-ordinated

**Figure 5. DRAGON performance over 569 cell cycle regulated genes from the Pramila and Orlando datasets.** (A) Scatterplot of the DRAGON estimated half-lives using two replicates taken from Pramila et al. dataset [18] (denoted by alpha30 and alpha38). (B) Scatterplot of the DRAGON estimated half-lives using two replicates taken from Orlando et al. dataset [19] (denoted by Orlando wtr1 and Orlando wtr2). (C) Histograms of the Pearson correlation values between time series relative to each gene in two replicates. Orlando et al. dataset shows a better consistency between replicates with respect to the Pramila et al. dataset. (D) Scatterplot of the DRAGON estimated half-lives using the Pramila et al. [18] and the Orlando et al. dataset [19]. The half-lives obtained by replicate datasets have been averaged. The half-lives estimated using Orlando dataset show slightly higher values with respect to those obtained using Pramila dataset, as shown by the deviation from the bisector line (dashed blue line). This is consistent with the slower cell cycle in Orlando experiment (2 hours) compared to that of Pramila experiment (1 hour).

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interplay between transcription and degradation. They have found
in yeast two basic regulatory strategies in response to stress. More
precisely, they measured changes of mRNA abundance and decay
rates in a yeast population subjected to oxidative and DNA
damage stress. By grouping genes according to the time point at
which the maximal (minimal) fold change is attained and
combining normalized (mean and variance) mRNA abundance
and decay rate data, they constructed a “stability versus folding”
(SF) diagram where change in mRNA stability relative to a
reference state (mean value in our case) is plotted against the
maximal fold change. Using yeast expression time-course data
obtained in response to an oxidative stress and a DNA damage,

![Figure 6. Periodic behavior of average half-life during the reproductive and metabolic cycle.](image)

Average DRAGON estimated messenger half-life values corresponding to genes having the same expression peak timing, indicated on the x-axis (dark blue points). Standard deviations are drawn as pale blue stripes. (A) yeast cell cycle using Pramila dataset, (B) yeast cell cycle using Orlando dataset and (C) yeast metabolic cycle. Strikingly, in all datasets the maximal average half-life is attained for genes induced during G2/M phase, including the metabolic cycle. The minimal average half-life, in all datasets, is attained at the G1/S transition phase. The bar charts show mean and standard deviation of half-lives during each stage.

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they were able to reveal two different strategies: a) a “counteracting regulation” strategy (see Figure 7A), characterized by genes in which an increase (decrease) in degradation rates counteracts a increase (decrease) in mRNA abundance, i.e. repressed genes are stabilized and induced genes are destabilized; b) a “synergistic regulation” strategy (see Figure 7B), characterized by genes in which an increase (decrease) in degradation rates is associated with a increase (decrease) in mRNA abundance, i.e. induced genes are stabilized and repressed genes are destabilized.

Shalem et al. also found that, progressing from early time points forward, the negative correlation (counteracting) was replaced with a positive correlation (synergistic). Such co-ordination strategy may permit crosstalk between different steps of mRNA biogenesis, providing a mechanism to control the order and timing of events [28]. The work of Shalem et al. has shown the importance of combining expression data with decay rates under the same experimental condition to reveal the underlying strategy of co-ordination of the two “regulatory arms”, namely transcription and degradation. Uncovering such relationships is certainly a fundamental task, since the underlying reciprocal influences between mRNA production and degradation are largely unexplored [10].

The DRAGON algorithm, by estimating half-lives directly from gene expression data under specific conditions, allows the computational integration of mRNA abundance and decay rates data, making this powerful combined analysis possible when experimentally measured half-lives are not available.

We computed SF diagrams for *P. falciparum* IDC, yeast cell cycle (Pramila et al. dataset) and metabolic cycle (shown in Figure 8). In panels A,C and E each blue dot corresponds to a Pearson correlation of the SF diagram at the peak time point indicated on the x-axis, for the three datasets. In panels B,D and F the SF diagrams corresponding to the correlation values indicated by the arrows in panels A,C and E are displayed. The arrows point to maximal negative (red dots in panels B,D,F and red arrows in panels A,C,E) and maximal positive correlation values (green dots in panels B,D,F and green arrows in panels A,C,E).

Figure 7. Illustration of the counteracting and synergistic regulatory strategy. The stability/folding diagrams (SF), introduced by Shalem et al. in [10], show the change in mRNA stability relative to the average value plotted against the maximal fold change. (A) Counteracting strategy (negative correlation): induced genes are destabilized and repressed genes are stabilized. (B) Synergistic strategy (positive correlation): induced genes are stabilized and repressed genes are destabilized. doi:10.1371/journal.pcbi.1002772.g007

Advantages and disadvantages of the method

The main advantage of the DRAGON algorithm consists in the estimation of the mRNA half-lives directly from gene expression time-course during condition-specific experiments. Moreover it estimates the correlation among promoter activities between pairs of genes. Another advantage of the algorithm lies in its robustness. Specifically, we observed that even if the accuracy of the absolute values of the estimated half-lives can be influenced by many factors (such as the number of points in the time series, the accuracy of the measurements, the time interval between samples, the choice of the thresholds for the outliers, etc.), the ranking of half-lives is insensitive to the factors mentioned above.

The main disadvantages are the following: DRAGON can work only with time-series under the same experimental condition and cannot handle steady-state values under different conditions. As a general rule a reliable estimate requires at least 10–12 time samples, i.e. a number significantly larger than the number of parameters to be estimated (this rule is not obviously always applicable as the required number of points depends strongly on the signal to noise ratio) and a sampling time not larger than the expected average half-life. If no information is available about the correlation of promoter activities, as a rule of thumb, a set of at least 50–100 time series must be processed together in order to have reliable half-lives estimates. One basic hypothesis is that the half-life of a transcript is approximately constant during the time course of the experiment, thus a substantial change of...
its value would yield an unreliable estimate. These problems can be handled by performing more measurements using a shorter sample time, or by considering moving time windows. The computational overhead can be significant: for a sample of 1000 time series there are $10^6$ pairs to analyze, requiring a computation of about 150 hours on a medium-speed single-processor machine capable of analyzing 2 pairs per second.

**Conclusion**

Our analysis supports and strengthens Shalem et al. conclusions about the coordination of transcriptional and mRNA degradation in the cell in response to stress. We have demonstrated that during periodic processes, such as the *P. falciparum* IDC, the reproductive cell cycle and the metabolic cycle, the alternative interplays between changes in mRNA stability and changes in mRNA abundance are activated by periodically switching from a counteracting to a synergistic regulation. In light of these results, the classical vision of periodic processes as the result of serial transcription factor sequential activation, should be re-considered from a broader point of view by including post-transcriptional regulation and coordination.

**Figure 8. Temporal progress on the stability/folding (SF) diagrams.** In all cases under study, we found a progressive shift from an inverse (counteracting) to a direct (synergistic) relationship. (A) *Plasmodium falciparum* IDC, (C) yeast cell cycle (Pramila et al. dataset) and (E) yeast metabolic cycle. In panels A,C and E each blue dot corresponds to a Pearson correlation of the SF diagram at the peak time point indicated on the x-axis, for the three datasets. In panels B,D and F the SF diagrams corresponding to the correlation values indicated by arrows in panels A,C and E are displayed. The arrows point to maximal negative (red dots in panels B,D,F and red arrows in panels A,C,E) and maximal positive correlation values (green dots in panels B,D,F and corresponding green arrows in panels A,C,E).

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Materials and Methods

Stochastic modeling of expression kinetics and Kalman filtering

We defined as \( x_i(t) \) the time profile of the expression of gene \( i \) at time \( t \). The underlying conservation equation simply stems from the observation that the rate of change of \( x_i(t) \) with time, i.e. its time derivative \( \dot{x}_i(t) \), must equal the difference between the production and degradation term. Based on experimental evidence [2], the degradation is well described by a first order term. The dynamics of the \( i \)-th transcript is therefore described by

\[
\dot{x}_i(t) = -k_i x_i(t) + P_i(t)
\]

where \( k_i \) is the mRNA decay rate of \( i \)-th messenger. This value is linked to the half-life \( h_i \) of the transcript by the relation \( h_i = \ln(2)/k_i \). \( P_i(t) \) is the \( i \)-th gene promoter activity regulated by transcription factors. Such regulation occurs by triggering or suppressing the transcription of the \( i \)-th gene, thus we have \( P_i(t) \geq 0 \). Moreover, the observed measure \( y_i(t) \) is also a noisy time-series, thus we have

\[
y_i(t) = x_i(t) + \sigma_{y_i} n_{y_i}(t)
\]

where \( \sigma_{y_i} \) is the standard deviation of measurements white noise \( n_{y_i} \) (see supplementary material Text S1 for an example of the identification procedure). We considered a generic pair of expression time profiles characterized by the presence of two terms: a stochastically correlated promoter activity \( u(t) \) and a gene-specific term \( n_{y_i}(t) \). We then considered the case:

\[
P_i(t) = b_i u(t) + \sigma_{v_i} n_{v_i}(t),
\]

where \( b_i \) is a scaling factor accounting for the relative contribution to the overall promoter activity regarding gene \( i \). The term \( \sigma_{v_i} n_{v_i} \) models the part of the promoter activity which is not common to the pair. We model this part by means of a noise term, \( n_{v_i} \) which is assumed to be a white noise. The common part \( u(t) \) is modeled as a Wiener process:

\[
\dot{u}(t) = \sigma_u n_u(t)
\]

where \( n_u(t) \) is white noise. Thus \( u(t) = u(0) + \int_0^t \sigma_u n_u(t) \, dt \). The complete mathematical dynamic model for two transcripts \( j \) and \( h \), together with their respective measurement equations, is

\[
\begin{align*}
\dot{x}_j(t) &= -k_j x_j(t) + b_j u(t) + \sigma_{y_j} n_{y_j}(t) \\
\dot{x}_h(t) &= -k_h x_h(t) + b_h u(t) + \sigma_{y_h} n_{y_h}(t) \\
\dot{u}(t) &= \sigma_u n_u(t) \\
y_j(t) &= x_j(t) + \sigma_{y_j} n_{y_j}(t) \\
y_h(t) &= x_h(t) + \sigma_{y_h} n_{y_h}(t)
\end{align*}
\]

We can rewrite the linear dynamic system (9) using a compact matrix notation

\[
\begin{align*}
x(t) &= [x_j(t) \quad x_h(t)] \\
y(t) &= [y_j(t) \quad y_h(t)]
\end{align*}
\]

where

\[
\begin{align*}
N_j(t) &= [n_{v_j}(t) \quad n_{v_h}(t)] \\
N_h(t) &= [n_{v_j}(t) \quad n_{v_h}(t)]
\end{align*}
\]

and

\[
\begin{align*}
A &= \begin{bmatrix} -k_j & 0 \\ 0 & -k_h \end{bmatrix}, \\
F &= \begin{bmatrix} \sigma_y & 0 \\ 0 & \sigma_y \end{bmatrix}, \\
G &= \begin{bmatrix} \sigma_y \\ 0 \end{bmatrix}, \\
C &= \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix},
\end{align*}
\]

Since the dynamic system (10) is linear, it can be exactly discretized [see [24]] for a given time interval \( \Delta t \), corresponding to the time interval between two consecutive measurements. The \( k \)-th measurements corresponds to \( t = k \Delta t \), thus in the discretized system we can use \( k \) in place of \( k \Delta t \), to keep the notation simple.

The solution of the linear dynamic system (10) is

\[
x(t) = e^{A(t-t_0)} x(t_0) + \int_{t_0}^t e^{A(t-t')} F(t') N_j(x(t')) \, dt'
\]

and its discretized form is

\[
\begin{align*}
x(k+1) &= A_d x(k) + \Psi_2 N_j(k) \\
y(k) &= C x(k) + G N_j(k)
\end{align*}
\]

where

\[
A_d = e^{A \Delta t},
\]

and \( \Psi \) is the covariance matrix defined by

\[
\Psi = \int_0^{\Delta t} e^{A \theta} F F^T e^{A^T \theta} \, d\theta.
\]

The unknown parameters of the model to be estimated are \( k_i, b_i, \sigma_{y_i}, \) with \( i = 1, 2 \), \( \sigma_u \) and \( \sigma_y \). The state variables of the system are \( x_j(t), x_h(t) \) and \( u(t) \). For each given choice of the parameters we used the Kalman filter [25] to estimate of the state variables.

The Kalman filter equation uses a feedback control strategy. It contains a prediction term for projecting forward (in time) the current state to obtain the a priori estimate, and a correcting term for incorporating a new measurement into the a priori estimate to obtain an improved a posteriori estimate

\[
\hat{x}(k+1) = A_d \hat{x}(k) + K_P(\theta) [y(k+1) - CA_d \hat{x}(k-1)]
\]

where \( K_P(\theta) \) is the prediction Kalman gain that depends on the parameters \( \theta \) of the stochastic equation.
For each choice of \( \theta \) we run the Kalman filter. A probability value is associated to the resulting estimation. These values measures the probability that the current parametrization of the model generates the measured time series. Denoting by \( \nu_s(k) = y(k) - C A_d x(k - 1) \) the innovation of the stochastic process, \( \nu_o(k) \) is a sequence of independent gaussian random variables with covariance \( \Psi_o = (C A_d^{-1} K P + I) G^2 \). The optimal set \( \theta \) of parameters if therefore chosen according to a maximum likelihood criterion as the choice corresponding to the maximum of the a priori probability density of the innovation sequence. This corresponds to the minimum of the likelihood function

\[
J(\theta) = n \ln(\det(\Psi_o)) + \sum_{k=1}^{n} \nu_o^T(k) \Psi_o^{-1} \nu_o(k).
\]

where \( n \) is the number of samples. We are interested in the half life \( h_i = \ln(2)/k_i \) of the \( i \)-th messenger. To use all the available information and make the method robust with respect to measurement and estimation errors, we have designed the following algorithm (see Supplementary Figure S6):

1. Given a set of \( N \) mRNA time profiles, perform the maximum likelihood estimation for every pair \( (i,j) \) and compute the corresponding \( k_i \) and \( k_j \).

2. For each pair \( (i,j) \) compute the ratio matrix \( M \) whose elements are the ratios between the half-lives of gene \( i \) and gene \( j \). The matrix \( M \) is generally not symmetric due to the presence of outliers and numerical sensitivity. Thus we defined the ratio estimation by row as \( m_{ij}^r = h_i / h_j \) and by column as \( m_{ij}^c = h_j / h_i \). The matrix \( M \) contains all the ratios \( h_i / h_j \) on the \( i \)-th row, and all the ratios \( h_j / h_i \) on the \( i \)-th column. Let us denote \( \Sigma_i^r \) the sum of the \( i \)-th row, \( \Sigma_i^c \) the sum of the \( i \)-th column, and \(< > \) the mean operator, that is,

\[
\Sigma_i^r = \sum_{j=1}^{N} h_j^r = \frac{1}{N} \sum_{j=1}^{N} h_j^r, \quad \Sigma_i^c = \sum_{j=1}^{N} h_j^c = \frac{1}{N} \sum_{j=1}^{N} h_j^c
\]

\[
< \Sigma_i > = \frac{1}{N} \sum_{j=1}^{N} h_j^r = \frac{1}{N} \sum_{j=1}^{N} h_j^c = \frac{1}{N} \sum_{j=1}^{N} h_j^r = \frac{1}{N} \sum_{j=1}^{N} h_j^c.
\]

3. Given \( M \), delete outliers to obtain a final matrix \( M^* \). First, compute the probability density (using a smoothing kernel approach) of all the entries \( m_{ij} \) and delete those values below a probability of \( 10^{-3} \) of occurring in the distribution. Second, since ideally \( m_{ij}^r m_{ij}^c = 1 \), we considered as outliers those pairs such that \( m_{ij}^r m_{ij}^c < 1 / \sqrt{2} \).

4. On the resulting \( M^* \) matrix compute for each transcript \( i \) two estimates of its half-life \( h_i^r \) and \( h_i^c \), using equations (15), (16) and (17). We obtained

\[
\hat{h}_i^r = \frac{\Sigma_i^r}{\Sigma_i^r}, \quad \hat{h}_i^c = \frac{\Sigma_i^c}{\Sigma_i^c}.
\]

This computation requires the value of \( \hat{h} \). When this value is known for the group of transcripts under analysis the measured value can be used. Otherwise, letting \( \hat{h} = 1 \) one can obtain half-life values that are relative to the average half-life of the group. However, we have followed a third approach. All the results reported in this paper have been obtained by replacing \( \hat{h} \) with the geometric mean of \( \hat{h}_i^r \) and \( \hat{h}_i^c \), that is

\[
\hat{h} = \left( \prod_{i=1}^{N} \hat{h}_i^r \prod_{i=1}^{N} \hat{h}_i^c \right)^{1/N}.
\]

The final estimate of the half-life \( \hat{h}_i \) for the \( i \)-th gene is computed as the weighted average of \( \hat{h}_i^r \) and \( \hat{h}_i^c \) using as weights the respective variances \( \sigma_{h_i}^2 \) and \( \sigma_{h_i}^2 \) as follows

\[
\hat{h}_i = \left( \frac{h_i^r}{\sigma_{h_i}^2} + \frac{h_i^c}{\sigma_{h_i}^2} \right)^{-1}
\]

where \( \sigma_{h_i} \) and \( \sigma_{h_i} \) are the standard deviation of the \( \hat{h}_i^r \) and \( \hat{h}_i^c \), respectively.

5. We considered as a quality index for each estimated half-life \( \hat{h}_i \)

\[
C_i = \frac{1}{N} \sum_{j=1}^{N} \frac{\Psi_{ij}}{\Psi_{ij} + \Psi_{ij}}
\]

where \( \Psi_{ij} \) is noise variances of the discrete system (12) and \( \Psi_{ij} \) (see equation (13)) is the mutual covariance of the state noise between time series \( x_i \) and \( x_i \). Thus, high values of \( C_i \) imply the presence of a correlation between \( P_i(t) \) and \( P_i(t) \) in equation (9). We removed the half-lives having a \( C_i \) value smaller than the 10th percentile of its distribution.

Datasets

Public experimental data used throughout the paper are described in Table 1 (experimental half-lives measurements) and in Table 2 (gene expression time series). Pramila et al. in [18] and Orlando et al. in [19] experimentally measured genome-wide gene expression data during the reproductive cell cycle. We considered the ranking provided by the combined test developed by de Lichtenberg et al. [29] for each replicate for the two datasets and, among the list of 1000 genes with highest ranking, we selected those common to all datasets. We ended up with a list of 569 genes that we used for half-life estimation. Tu et al. in [20] experimentally measured genome-wide gene expression data during the metabolic cell cycle. We selected 1000 genes with the best periodicity score according to [20]. Of the 1000 genes, DRAGON estimated half-lives are 939.
**Data processing**

**Half-lives determination of genes induced during each stage of* P. falciparum* IDC.** Shock *et al.* in [17] experimentally measured genome-wide values of decay rates for each gene in each of the four stages of the IDC. To obtain a single half-life value for each messenger, we performed a k-means clustering of microarray gene expression data [16] by considering 3 stages (according to [16]: early ring, ring, trophozoite, schizont and late schizont). Then, we merged the early ring cluster with the ring cluster to obtain the same stages as in Shock *et al.* Among the 4408 genes in [16] we chose 1000 genes with the best periodicity score (power signal/power total ratio) according to [16]. Of the 1000 genes, DRAGON estimated half-lives are 967, available experimental half-lives are 675. Both data are available over a set of 616 genes.

**Expression peak timing estimation**

To estimate peak timing, for a given noisy gene expression time profile, we preliminary performed the smoothing algorithm presented by Bar-Joseph *et al.* in [30]. The algorithm employs two parameters: grid \( g \) (number of spline curves) and classes \( c \) (number of classes to use for clustering). In particular, for Pramila datasets we used \( g = 4 \) and \( c = 2 \), for Orlando datasets we used \( g = 8 \) and \( c = 2 \), for Tu dataset we used \( g = 10 \) and \( c = 5 \), for Malaria dataset we used \( g = 10 \) and \( c = 4 \).

**Half-life estimations**

DRAGON estimated half-lives are provided as supplementary materials Tables S1, S2, S3, S4, S5, S6, S7, described Supporting Information section.

Matlab code will be provided upon request.

Additional data and information can be found at web site http://www.dis.uniroma1.it/~farina/dragon.

**Supporting Information**

**Figure S1** Functional categories analysis in yeast *S. cerevisiae* during asynchronous growth measured by three laboratories. Three genome-wide studies are considered: Grigull *et al.*, Wang *et al.* and Munchel *et al.* (A) Average mRNA half-lives in both studies Wang *et al.* and the Grigull *et al.* datasets for 111 functional categories from the yeast GO Biological Process database (http://www.geneontology.org) that are represented in the set of 2863 transcripts by 5 or more members. (B) compare, in the same way, the Munchel *et al.* and the Wang *et al.* datasets.

**Figure S2** Kinetics of gene induction. Panels A–B show in silico experiments to illustrate some basic features of gene induction kinetics. The reference time profile with unity steady-state is plotted in black. The “ON” and “OFF” regions correspond to the turning “ON” or “OFF” of the promoter activity. (A) Induction kinetic of transcripts having the same half-life value and, as a consequence, the same speed of response. The higher (or lower) steady-state value of the red and blue time profiles is due only to an increased (or decreased) transcription rate. (B) Induction kinetic of transcripts having different half-lives. The time profile plotted in red corresponds to an unstable transcript. It has a faster induction and relaxation profile but a lower steady-state value. By contrast, the blue one has a higher half-life value, resulting in a higher steady state value but a slower response. The example illustrates that, to obtain both a fast response and an high steady-state value, the regulatory strategy must destabilize transcriptionally up-regulated genes.

**Figure S3** Functional categories analysis for DRAGON estimations using *P. falciparum* IDC data. Average mRNA half-lives in both studies, DRAGON estimations versus and experimentally measured by Shock *et al.* half-lives, for 12 functional categories from the *P. falciparum* GO annotation database (http://www.geneontology.org) that are represented in the set of 616 transcripts by 5 or more members.

**Figure S4** GO annotations of genes with extreme half-lives in *S. cerevisiae* DNA replication timing requires first the formation of the replication fork, then the production of the needed histones for chromatin assembling: such temporal sequence of events is consistent with a rapid turnover of the replication complex genes and a slow turnover of the histone genes.

**Figure S5** GO annotations of genes with extreme half-lives in *S. cerevisiae* Among unstable genes we also found the G1/S transition cyclins and among stable ones we found G2/M transition cyclins. In this case, the temporal sequence of events is the progression of the cell cycle from DNA replication to mitosis.

**Figure S6** DRAGON algorithm pipeline.

**Table S1** DRAGON estimated half-lives using alpha30 dataset. (XLSX)

**Table S2** DRAGON estimated half-lives using alpha38 dataset. (XLSX)

**Table S3** DRAGON estimated half-lives using Orlando replicate 1 dataset. (XLSX)

**Table S4** DRAGON estimated half-lives using Orlando replicate 2 dataset. (XLSX)

**Table S5** DRAGON estimated half-lives using metabolic cycle dataset. (XLSX)

**Table S6** DRAGON estimated half-lives using *P. falciparum* dataset. (XLSX)

**Table S7** DRAGON estimated normalized half-lives using all yeast datasets. (XLSX)

**Text S1** Example of parameter estimation through the Kalman filter.

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**Author Contributions**

Conceived and designed the experiments: LF AG. Performed the experiments: FC PP VC. Analyzed the data: PP LF. Wrote the paper: FC PP AG LF. Made the figures: LF.
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