Quasars as standard candles

III. Validation of a new sample for cosmological studies*

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ABSTRACT

We present a new catalogue of ∼2400 optically selected quasars with spectroscopic redshifts and X-ray observations from either Chandra or XMM–Newton. The sample can be used to investigate the non-linear relation between the ultraviolet (UV) and X-ray luminosity of quasars as well as to build a Hubble diagram up to a redshift of z ∼ 7.5. We selected sources that are neither reddened by dust in the optical and UV nor obscured by gas in the X-rays, and whose X-ray fluxes are free from flux-limit-related biases. After checking for any possible systematics, we confirm, in agreement with our previous works, that the X-ray to UV relation provides distance estimates matching those from supernovae up to z ∼ 1.5, and its slope shows no redshift evolution up to z ∼ 5. We provide a full description of the methodology for testing cosmological models, further supporting a trend whereby the Hubble diagram of quasars is well reproduced by the standard flat cold dark matter model up to z ∼ 1.5–2, but strong deviations emerge at higher redshifts. Since we have minimised all non-negligible systematic effects and proven the stability of the LX − LUV relation at high redshifts, we conclude that an evolution of the expansion rate of the Universe should be considered as a possible explanation for the observed deviation, rather than some systematic (redshift-dependent) effect associated with high-redshift quasars.

Key words. galaxies: active – quasars: general – quasars: supermassive black holes – methods: statistical

1. Introduction

Quasars are the most luminous and persistent energy sources in our Universe. As they can be observed up to a redshift of ∼7.5 (Bahados et al. 2018), when the age of the Universe was less than ∼700 million years old, quasars bear an extraordinary potential as cosmological probes. Several techniques making use of empirical correlations between quasar properties have been proposed to determine cosmological parameters such as the dark matter (ΩM) and dark energy (ΩΛ) content of the Universe. Examples include the relation between the continuum luminosity and the emission-line equivalent width (Baldwin 1977), or with the radius of the quasar broad-line region established via reverberation mapping (Watson et al. 2011). Another luminosity distance estimator combines the correlation between the quasar X-ray variability amplitude and its black hole (BH) mass (La Franca et al. 2014). Nonetheless, these correlations are affected by too large a dispersion (up to ≈0.6 dex), and they are typically applicable over a limited redshift range. Other promising methods employ geometric distances through, again, the broad-line region radius via reverberation mapping (Elvis & Karovska 2002), the luminosity of super-Eddington accreting quasars (Wang et al. 2013), the eigenvector formalism in the quasar main sequence (Marziani & Sulentic 2014), or the combination of spectroastrometry and reverberation mapping (Wang et al. 2020). All of these techniques still need some level of refinement, and (or) much higher sample statistics, in order to be competitive as cosmological tools.

Since 2015, our group has been developing a new technique that hinges on the observed non-linear relation between the ultraviolet (UV, at 2500 Å, LUV) and the X-ray (at 2 keV, LX) emission in quasars (e.g. Tananbaum et al. 1979, LX ∝ LUVγ, with γ ≈ 0.6) to provide an independent measurement of their distances,
thus turning quasars into standardizable candles and extending the distance modulus–redshift relation (or the so-called Hubble-Lemaître diagram) of supernovae Ia to a redshift range that is still poorly explored \((z > 2)\) \cite{RisalitiLusso2015}.

The applicability of our technique is based on the understanding that most of the observed dispersion in the \(L_X - L_{UV}\) relation is not intrinsic to the relation itself but due to observational issues, such as X-ray absorption by gas, UV extinction by dust, calibration uncertainties in the X-rays \cite{Lusso2019}, variability, and selection biases associated with the flux limits of the different samples. In fact, with an optimal selection of clean sources (i.e., where we can measure the intrinsic UV and X-ray quasar emission), the dispersion drops from 0.4 to \(\pm 0.2\) dex \cite{LussoRisaliti2016, LussoRisaliti2017, Arcodia2019}.

A key consequence of this technique is that the \(L_X - L_{UV}\) relation must be the manifestation of a universal mechanism at work in the quasar engines. However, the details on the physical process at the origin of this relation are still unknown \citep[e.g.][]{Haardt1991, Haardt1993, Nicastro2000, Merloni2003, LussoRisaliti2017, Arcodia2019}.

The main results of our work are that the distance modulus-redshift relation of quasars at \(z < 1.4\) is in agreement with that of supernovae Ia and with the concordance cold dark matter (ΛCDM) model \citep{RisalitiLusso2015, Lusso2019, Lusso2019}, yet a deviation from ΛCDM emerges at a higher redshift, with a statistical significance of about \(4\sigma\). If we interpret the latter result by considering an evolution of the dark energy equation of state in the form \(w(z) = w_0 + w_a z^2 / (1 + z)^2\), the data suggest that the dark energy density is increasing with time \citep{RisalitiLusso2019, Lusso2019}.

As our approach may still have shortcomings, we need to demonstrate that the observed deviation from ΛCDM at a redshift \(> 2\) is neither driven by systematics in the quasar sample selection nor by the procedure adopted to fit the quasar Hubble-Lemaître diagram. To build a quasar sample that can be utilised for cosmological purposes, both X-ray and UV data are required to cover the rest-frame 2 keV and 2500 Å. The choice of these monochromatic luminosities is rather arbitrary, and mostly based on historical reasons. It is possible that the \(L_X - L_{UV}\) relation is tighter with a different choice of the indicators of UV and X-ray emission \citep[see e.g.][]{Young2010}. A careful analysis of this issue may also provide new insights as to the physical process responsible for this relation. In the present analysis, we do not investigate the possible alternatives, but we focus on demonstrating that the commonly used relation is calibrated in a robust way and can be safely used as a tool to determine quasar distances.

In this third paper of our series, we thus mainly concentrate on the quasar sample selection, whilst we defer a detailed analysis of the cosmographic fitting technique we adopted to a forthcoming publication.

At the time of writing this paper, the most extended spectroscopic coverage in the UV is given by the Sloan Digital Sky Survey \citep{Parris2018}, which supplies more than \(\sim 500,000\) quasars with spectroscopic redshift up to \(z \sim 7\). This sample needs to be cross-matched with the current X-ray catalogues, namely the Chandra X-ray Catalogue \citep[CXC2.0,][]{Evans2010} and the 4XMM Data Release 9 \citep{Webb2020}, which contain all the X-ray sources detected by the Chandra and XMM–Newton observatories that are publicly available in the respective archives. These data need to be complemented by dedicated pointed observations of active galactic nuclei\(^1\) (AGN) at both low \((z < 0.1)\) and high \((z > 3)\) redshifts to extend the coverage and increase the sample statistics in the distance modulus–redshift relation.

The main aims of this manuscript are to discuss in detail all the criteria required to select a homogeneous sample of quasars for cosmological purposes from the above archives, and to present the key steps in fitting the Hubble-Lemaître diagram that can be adopted to reproduce our results. Specifically, we examine the effects on the sample selection and on the UV and X-ray flux measurements of dust extinction and host-galaxy contamination, gas absorption in the X-rays, and Eddington bias. We identify the quasars that can be used for testing cosmological models, and investigate in depth all the possible systematics in the quasar Hubble-Lemaître diagram as a function of the contaminants mentioned above.

Since a detailed spectroscopic UV and X-ray analysis can only be carried out only for a relatively small number of sources, our latest quasar sample presented here still relies also on broadband photometry in both UV and X-rays, as most of the parameters currently employed in our works, such as monochromatic UV and X-ray fluxes, UV colours and X-ray slopes, are derived from the photometric spectral energy distribution (SED) of our sources. In the future, we plan to gradually move towards a full spectroscopic analysis, as spectroscopy can deliver cleaner measurements of the relevant parameters.

The paper is constructed as follows. In Sect. 2 we present the different data sets employed to build the main quasar sample and the procedures adopted to measure the UV fluxes and slopes from the photometry. Section 3 discusses how the photometric quasar SEDs are constructed, whilst in Sect. 4 we justify the use of the rest-frame 2500 Å monochromatic fluxes derived from photometry. Section 5 is devoted to the presentation of our technique to compute the monochromatic X-ray emission and the photon index from the catalogued broadband flux values. Section 6 gives reasons for the analysis of the Hubble diagram where non-detections are neglected. In Sect. 7 we discuss the several quality filters employed to select a homogeneous sample of quasars for cosmological purposes, whilst in Sect. 8 we verify that the \(L_X - L_{UV}\) relation for the final “best” quasar sample does not evolve with redshift. Section 9 presents the main steps adopted to fit the Hubble diagram using a model independent technique (i.e., cosmography), whilst in Sect. 10 we fit the Hubble diagram with the most commonly used ΛCDM extension to test our fitting technique and to verify how different choices regarding the fitting method and the quasar sample selection affect the final results. In Sect. 11 we carry out an in-depth investigation on possible remaining systematics in the residuals of the quasar Hubble diagram, as a function of the parameters involved in the selection of the sample. Finally, we summarise our work and main results in Sect. 12. Although we mainly use fluxes, whenever luminosity values are reported we have assumed a standard flat ΛCDM cosmology with \(\Omega_M = 0.3\) and \(H_0 = 70\) km s\(^{-1}\) Mpc\(^{-1}\).

1 In the following we refer to AGN and quasars indistinctly.

2. The data set

The broad-line quasar sample we considered for our analysis has been assembled by combining seven different samples from both the literature and the public archives. The former group includes the samples at \(z = 3.0–3.3\) by \cite{Nardini2019}, \(4 < z < 7\) by \cite{Salvestrini2019}, \(z > 6\) by \cite{Vito2019}, and the XMM–XXL North quasar sample published by \cite{Menzel2016}. We then complemented this collection by including quasars from a cross-match of optical (i.e. the Sloan Digital Sky Survey) and X-ray public catalogues (i.e. XMM–Newton
and Chandra), which we refer to as SDSS–4XMM and SDSS–Chandra samples hereafter. Finally, we also added a local subset of AGN with UV (i.e. International Ultraviolet Explorer) data and X-ray archival information. The same order in which these samples have been introduced above is adopted as an order of priority to take into account all the possible overlaps. X-ray fluxes coming from pointed observations and medium-deep surveys (i.e. XMM–XXL) have been considered first, as they are generally more reliable. The main parent sample is composed by ~19000 objects spanning the redshift range 0.009 < z < 7.52. In Fig. 1 we present the distribution of rest-frame 2500 Å luminosities as a function of redshift for all the different quasar sub-samples. A summary of the sample statistics is provided in Table 1, whilst below we describe in detail of how each sub-sample has been constructed.

2.1. The SDSS–4XMM sample

The bulk of the data belongs to the Sloan Digital Sky Survey quasar catalogue, Data Release 14 (hereafter SDSS–DR14; Pâris et al. 2018). The catalogue contains 526356 optically selected quasars detected over 9376 deg$^2$ with robust identification and spectroscopic redshift. Firstly, we removed from the catalogue all quasars flagged as broad absorption lines (BALs), where sources with the BI_CIV=0 flag are non-BALs) and kept all the objects with a measurement in all the SDSS magnitudes. This preliminary selection leads to 503744 quasars.

We note that the BAL classification in the SDSS–DR14 quasar catalogue is based on a fully automated detection procedure on CIV absorption troughs for sources at z > 1.57. Hence, a number of BAL quasars might still be included in this preliminary sample. BAL quasars are often found in galaxies with red optical and UV colours and hard X-ray spectra (e.g. Gallagher et al. 2006), the latter suggesting that their relative X-ray weakness could be primarily due to gas absorption. The selection criteria discussed in Sect. 7 efficiently remove red and X-ray absorbed quasars, possibly excluding most unclassified BALs from the final sample.

The photometric rest-frame spectral energy distribution (SED), whose derivation is discussed in Sect. 3, is then used to define the parameters required to exclude radio-loud, dust-absorbed or host-galaxy contaminated sources. Also the rest-frame monochromatic luminosities at 2500 Å are obtained from the photometric SEDs. For comparison purposes, as in our previous works on this topic, we discard bright radio quasars through the same radio loudness parameter, R, as that used in Shen et al. (2011), which is defined as the ratio of the rest-frame luminosities at 6 cm and 2500 Å (i.e. R = L$_{6cm}$/L$_{2500\AA}$). A quasar is then classified as radio-loud if R > 10. We computed R for the 17561 objects with a FIRST detection, 16315 of which are indeed radio-loud and have been therefore excluded from our sample.

To further remove powerful radio-loud quasars we considered the catalogue published by Mingo et al. (2016), which is currently the largest available Mid-Infrared (WISE), X-ray (3XMM) and Radio (FIRST+NVSS) collection (MIXR) of AGN and star-forming galaxies: 2753 sources, 918 of which are considered radio-loud based on multwavelength diagnostics (we refer to their paper for details). We excluded 349 quasars in our sample flagged as radio-loud in the MIXR catalogue within a matching radius of 3 arcsec. This yields 487080 SDSS radio-quiet quasars with a L$_{UV}$ measurement.

This SDSS quasar sample is then cross-matched with the latest XMM–Newton source catalogue 4XMM–DR9 (Webb et al. 2020). 4XMM–DR9 is the fourth generation catalogue of serendipitous X-ray sources, which contains 810795 detections (550124 unique X-ray sources) made publicly available by 2018 December 18. The net sky area covered (taking into account overlaps between observations) is ~1152 deg$^2$, for a net exposure time of ≥1 ksec.

To select reliable X-ray detections, we have applied the following quality cuts in the 4XMM–DR9 catalogue: SUM_FLAG < 3 (low level of spurious detections), OBJ_CLASS ≤3

Table 1. Summary of sample statistics.

| Sample               | Initial (1) | Main (2) | Selected (3) | Reference |
|----------------------|-------------|----------|--------------|-----------|
| XMM–Newton z ≃ 3    | 29          | 29       | 14           | 1         |
| XMM–Newton z ≃ 4    | 1           | 1        | 1            | 2         |
| High-z               | 64          | 64       | 35           | 3         |
| XXL                  | 840         | 840      | 106          | 4         |
| SDSS–4XMM            | 13800       | 9252     | 1644         | 5         |
| SDSS–Chandra         | 7036        | 2392     | 608          | 6         |
| Local AGN            | 15          | 15       | 13           | 7         |
| Total                | 21785       | 12593    | 2421         |           |

Notes. (1) These number counts refer to the sample statistics before correcting for overlaps amongst the sub-samples. (2) Sample statistics after accounting for overlaps and the quality pre-selection described in Sect. 7. The order of priority decreases from the top to the bottom row. (3) Sample statistics in the final cleaned quasar sample after the various filtering steps; see Sect. 7 for details.

References. Various samples: 1: Nardini et al. (2019); 2: see Sect. 2.5; 3: Salvestrini et al. (2019) and Vito et al. (2019); 4: Menzel et al. (2016); 5: see Sect. 2.1; 6: Bisogni et al. (in prep.); 7: Sect. 2.7.

http://xmmssc.irap.omp.eu/Catalogue/4XMM-DR9/4XMM-DR9_Catalogue_User_Guide.html

Fig. 1. Distribution of luminosities at rest-frame 2500 Å as a function of redshift for the main (grey points) and the selected (cleaned) samples (see Sect. 7.4). Brown and yellow squares: high-z sample (Salvestrini et al. 2019; Vito et al. 2019, see Sect. 2.6), cyan points: SDSS–4XMM (Sec. 2.1), brown triangles: XMM–XXL (Menzel et al. 2016, Sect. 2.3), orange pentagons: the local AGN sample (Sect. 2.7), red stars: z ≃ 3 quasar sample (Nardini et al. 2019, green star: new z ≃ 4 quasar (Sect. 2.5), gold pentagons: SDSS–Chandra (Sect. 2.2).
(quality classification of the whole observation) and EP_TIME > 0 (EPIC exposure time available). These filters lead to 692 815 X-ray detections. We have adopted a maximum separation of 3 arcsec to provide optical classification and spectroscopic redshift for all the cross-matched objects. This yields 22 196 XMM–Newton observations; 13 858 unique sources (3976 of which have ≥ 2 observations) covering the redshift range 0.056 < z < 4.306.

Following the results presented by Lusso & Risaliti (2016, LR16 hereafter; see their Sect. 4), we decided to average all X-ray observations for sources with multiple detections that meet our selection cuts, including that associated with the Eddington bias (see Sect. 7.3 for details). In this way, we reduce the effect of X-ray variability on the dispersion (~0.12 dex, see Sect. 4 in LR16) by using only unbiased detections.

For each XMM–Newton observation, we have computed the EPIC sensitivity (5σ minimum detectable flux) at 2 keV following a similar approach as in LR16. We first estimated the minimum detectable flux in the soft band for both pn and MOS as a function of the on-time exposure following the relations plotted in Fig. 3 by Watson et al. (2001). The total MOS on-time exposure is the one with the largest exposure value between the two individual cameras, MOS1 and MOS2. We then corrected this sensitivity for the pn and MOS, using the same vignetting correction for both cameras at the energy of 1.5 keV, as a function of their respective off-axis angles, where the smaller value between the two individual cameras is again assumed for the MOS. The sensitivity at 2 keV (F_min) is then estimated for both pn and MOS, assuming a photon index of 1.7, following the same approach as in LR16. Finally, we have prioritised the pn sensitivity flux values over the MOS when available.

2.2. The SDSS–Chandra sample

To further increase the statistics, we also cross-matched the SDSS–DR14 quasar catalogue with the second release of the Chandra Source Catalog (CSC2.0). The CSC2.0 (Evans et al. 2010) contains ∼315 000 X-ray sources observed in 10 382 Chandra ACIS-I and HRC-I imaging observations publicly released prior to 2015. A cross-match of these two catalogues, with a matching radius of 3 arcsec, leads to 7036 unique objects. The detailed analysis of this sample will be presented in a forthcoming publication (Bisogni et al., in prep.). Briefly, from this sample we excluded radio-loud and BAL quasars following the same approach adopted with the SDSS–4XMM sample. SEDs were also compiled for all the quasars (see Sect. 3), which were then used to estimate both L_{UV} and optical colours and thus select objects with low levels of dust reddening and host-galaxy contamination.

CSC2.0 provides photometric information and data products for each source, already reduced and ready to use for spectroscopic analysis. We selected all the AGN with at least one measurement of the flux in the soft band and with an off-axis angle <10 arcmin (3569 quasar observations, 2392 single quasars). We performed a full spectral analysis with the XSPEC v12.10.1b X-ray spectral fitting package (Arnaud 1996). For each observation, we assumed a model consisting of a power law with Galactic absorption, as provided by the CSC 2.0 catalogue at the source location. The spectral analysis provides us with the rest-frame flux at 2 keV and its uncertainty. Overall, the Chandra data have a reasonable signal-to-noise ratio (S/N > 5 in the soft band) that ensures uncertainties on F_{X} of the order of 0.15 dex or better.

Flux limits are estimated for any given Chandra observation by computing the percentage of net counts to deduce the significance of the source detection, and a factor that takes into account the level of background, P_{bg}. The 0.5–2 keV and 2–7 keV fluxes are then multiplied by P_{bg} to obtain an approximate value of the background flux in these energy bands. The flux limit in each energy band is then estimated from the background flux by assuming a minimum signal-to-noise ratio of 3. The flux limit at 2 keV is finally inferred by interpolation (or extrapolation) of the band flux limits, depending on the redshift of the source.

2.3. The XMM–XXL sample

We also considered the AGN sample published by Menzel et al. (2016) from the equatorial subregion of the XMM–Newton XXL (XMM–XXL, PI: Pierre), that is, XMM–XXL North (in the following we refer to this sample as XXL for simplicity), which overlaps with the SDSS–DR8 imaging survey, XMM–XXL North is a medium-depth (10 ks per pointing) X-ray survey distributed around the area of the original 11 deg² XMM–LSS survey. The total catalogue contains 2570 X-ray AGN with optical counterparts, spectroscopic redshifts and emission lines information. From the main sample, we considered only the AGN classified as (point-like) optically unobscured (flagged as BLAGN; 1353 sources). To have consistent measurements of optical and UV luminosities and redshifts amongst the different samples, we cross-matched the XXL BLAGN with the SDSS–DR14 quasar catalogue (with 3 arcsec matching radius) finding 1067 objects. We have then included only the AGN with available SDSS photometry and classified as non-BAL, leading to 915 AGN. For this sample, we compiled the photometric SEDs following the same approach as in Sect. 3, and computed luminosities at various rest-frame wavelengths (e.g. 2500 Å, 1450 Å, 6 cm), optical and UV colours (Γ₁, Γ₂, see Sect. 7.1 for details) and radio loudness. The latter parameter further excludes 75 AGN, leading to a final sample of 840 sources.

2.4. The z ∼ 3 quasar sample

We included a sample of 29 bright (L_{bol} > 10^{47} erg s^{-1}) quasars at z ≳ 3 with X-ray observations obtained from an extensive campaign performed with XMM–Newton (cycle 16, proposal ID: 080395, PI: Risaliti). This campaign targeted 30 quasars in the z = 3.0–3.3 redshift range for a total exposure of 1.13 Ms. This sample, selected in the optical from the SDSS Data Release 7 to be representative of the most luminous, intrinsically blue quasar population, boasts by construction a remarkable degree of homogeneity in terms of optical and UV properties. The X-ray data have been extensively analysed: the interested reader should refer to Nardini et al. (2019) for details.

2.5. New z ∼ 4 quasar

We also included one new optically-selected SDSS quasar at z = 4.109, J074711.14+273903.3, whose X-ray observation was

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3 For more details the reader should refer to the 4XMM catalogue user guide.
4 The total good (after flares removal) exposure time (in seconds) of the CCD over which the source is detected.
5 http://cxc.harvard.edu/csc2/
6 All the X-ray info can be downloaded from the CSCview application http://cda.harvard.edu/cscview/
7 One quasar in this sample turned out to be radio-loud, although not flagged as such in the SDSS–DR7 catalogue, so we exclude it from the present analysis.
obtained as part of a proposed large programme with XMM–Newton (cycle 18, proposal ID: 084497, PI: Lusso). This is the only target actually observed from its parent sample, which consisted of 19 quasars in the $z \approx 4$ redshift range for a total exposure of 1.34 Ms. The sample had been selected in the optical from the SDSS Data Release 14 with the same criteria described in Sect. 2.4. The X-ray spectrum of this quasar has been analysed following the same procedure as presented by Nardini et al. (2019), and we decided to include this source in the current sample as a proof of feasibility for future campaigns.

### 2.6. The high redshift sample

To improve the coverage at high redshifts, we considered two additional samples of $z > 4$ quasars with pointed X-ray observations published by Salvestrini et al. (2019) and Vito et al. (2019).

The Salvestrini et al. (2019) quasar sample consists of 53 objects in the redshift range 4.01 < $z$ < 7.08, which benefit from a moderate-quality coverage in the UV and X-ray energy bands. Of the 53 quasars, 47 objects were observed with Chandra and 9 with XMM–Newton. The galaxies ULAS J1120+0641, SDSS J114816.7+525150.4 and SDSS 1030+0524 have been observed by both satellites. The authors performed a full X-ray spectral analysis of the archival data, we thus refer to their paper for details. The majority of the quasars in this sample (33 out of 53) have $L_{\text{UV}}$ measurements from the SDSS–DR7 quasar catalogue. For the remaining quasars, the $L_{\text{UV}}$ values are computed by extrapolating the UV spectra to longer wavelengths with a fixed continuum slope (see their Sect. 4 and Appendix B for further details).

Vito et al. (2019) published a sample of 25 quasars at $z > 6$ with either archival data (15 objects) or new Chandra observations (10 sources), which were selected to have virial black-hole mass estimates from Mg II line spectroscopy. All the X-ray data were reprocessed by the authors (see their Sect. 3.1), whilst the $L_{\text{UV}}$ values were computed from the 1450 Å magnitude assuming a power-law spectrum ($F_\nu \propto \nu^\alpha$) with $\alpha = −0.3$ (see their Sect. 4.1). We excluded from their sample 3 BAL candidates, 1 weak line quasar, all sources with an upper limit in $\alpha_{\text{Fe}}$ (i.e. X-ray undetected) and all radio-loud sources, for a total of 9 quasars. For the remaining 16 sources, we found five overlaps with the Salvestrini et al. (2019) sample, so the final number of quasars included from Vito et al. (2019) is 11 sources. This sample also contains the highest-redshift quasar observed so far, ULAS J134208.10+092838.61, at $z = 7.54$ (Baños et al. 2018).

### 2.7. The local quasar sample

To anchor the normalisation of the quasar Hubble diagram with supernovae Ia, we need to extend the coverage at very low redshifts (0.009 < $z$ < 0.1). We searched for all the local AGN with ultraviolet data from the International Ultraviolet Explorer (IUE) in the Mikulski Archive for Space Telescopes (MAST). We chose to use the reduced spectra from the long-wavelength prime (LWP) camera of IUE, which spans the wavelength interval 1845–2980 Å, thus always covering the rest-frame 2500 Å at the redshifts of interest. We then considered all AGN with X-ray data available in the XMM–Newton archive or the with X-ray flux values in the literature, finding 17 objects, 11 of which with $\geq$ UV spectra (although the majority consists of consecutive observations). In this sample, NGC 1566 and NGC 7603 are well known highly variable and changing look sources, so we excluded them from the starting sample. Multiple UV spectra for the remaining AGN have been stacked, verifying that the inclusion of non consecutive observations does not change the final composite for each AGN.

We then carried out a detailed spectral fitting of all the UV spectra using the publicly available QSFIt package (Calderone et al. 2017). We modelled each spectrum as follows: the Mg II emission line is reproduced by a combination of a broad (with a full-width at half-maximum, FWHM, larger than 2000 km s$^{-1}$) and a narrow (FWHM < 2000 km s$^{-1}$) profile, whilst the continuum includes the contributions from the host galaxy, the iron complex, the Balmer continuum and the AGN continuum. Spectra are also corrected for Galactic extinction using the parametrisation by Cardelli et al. (1989) and O’Donnell (1994), with a total selective extinction $A(V) / E(B−V) = 3.1$ (Calderone et al. 2017). The rest-frame 2500 Å luminosity is finally measured from the AGN continuum component only. An example of a UV spectral fit on one of the objects in the local AGN sample is shown in Fig. 2.

The X-ray information (soft and hard fluxes, photon index) has been taken from the literature. Most of the sources in the
Table 2. Properties of the local AGN sample.

| Name      | z        | log $F_{\text{UV}}$ | log $F_X$ | Ref. |
|-----------|----------|---------------------|-----------|------|
| Ark 120   | 0.0327   | $-24.88 \pm 0.01$   | $-28.37 \pm 0.002$ | 1    |
| Mkn 841   | 0.0364   | $-25.46 \pm 0.01$   | $-29.15 \pm 0.006$ | 1    |
| NGC 4593  | 0.0090   | $-25.79 \pm 0.01$   | $-28.38 \pm 0.002$ | 2    |
| HE 1029–1401 | 0.0858 | $-25.10 \pm 0.01$   | $-28.77 \pm 0.005$ | 1    |
| ESO 141–G055 | 0.0371 | $-25.10 \pm 0.01$   | $-28.43 \pm 0.005$ | 2    |
| IRAS 09149–6206 | 0.0573 | $-25.19 \pm 0.01$   | $-28.86 \pm 0.011$ | 3    |
| HE 1143–181 | 0.0329 | $-25.37 \pm 0.01$   | $-28.57 \pm 0.005$ | 1    |
| NGC 7469  | 0.0163   | $-25.02 \pm 0.01$   | $-28.55 \pm 0.003$ | 1    |
| Mkn 205   | 0.0708   | $-27.03 \pm 0.01$   | $-29.33 \pm 0.016$ | 1    |
| Mkn 926   | 0.0469   | $-25.25 \pm 0.02$   | $-28.54 \pm 0.007$ | 1    |
| Fairall 9 | 0.0470   | $-25.34 \pm 0.01$   | $-29.01 \pm 0.016$ | 1    |
| Mkn 1383  | 0.0866   | $-25.28 \pm 0.01$   | $-29.02 \pm 0.017$ | 1    |
| Mkn 509   | 0.0344   | $-24.89 \pm 0.01$   | $-28.43 \pm 0.003$ | 1    |
| Mkn 478   | 0.0791   | $-25.52 \pm 0.01$   | $-29.51 \pm 0.023$ | 1    |
| Mkn 352   | 0.0149   | $-26.84 \pm 0.01$   | $-29.03 \pm 0.011$ | 1    |

Notes. UV and X-ray fluxes are in units of log(erg s$^{-1}$ cm$^{-2}$ Hz$^{-1}$).

References. X-ray data: 1: Bianchi et al. (2009); 2: de Marco et al. (2009); 3: Ricci et al. (2017).

Concerning the former, the 2500 Å monochromatic flux has been adopted since the first studies on the topic, yet its choice was mainly based upon observational considerations. Indeed, the rest-frame 2500 Å at $z \approx 1$ corresponds to the observed $V$ band, which was available for a significant number of sources. Additionally, this rest-frame UV wavelength is less affected by host-galaxy contamination (dominant at low luminosity for $L_\lambda > 4000$ Å) and intergalactic absorption ($L_\lambda < 1450$ Å) than other regions of the AGN SED, thus representing the ideal proxy of the intrinsic disc emission. This notwithstanding, the photometric 2500 Å flux might be contaminated by strong Fe II line emission (see Fig. 2), which can introduce systematic uncertainties on the photometrically derived $F_{\text{UV}}$ values (e.g. Netzer 2019).

Rakshit et al. (2020) recently published spectroscopic measurements for more than 500 000 quasars selected from the SDSS–DR14 quasar catalogue. They performed a homogeneous analysis of the SDSS spectra to estimate the continuum and line properties (e.g. $H\alpha$, $H\beta$, Mg II, C IV, and Ly $\alpha$) of these sources. This catalogue also provides a measurement of the 3000 Å luminosity, which is the closest wavelength to the one of our interest. We have therefore estimated the 3000 Å monochromatic luminosities from the photometric SEDs for all the SDSS–4XMM quasars in the initial sample of 13 800 quasars, similarly to what we have done at 2500 Å. Figure 4 shows the comparison...
A dispersion around the mean of 0.18 dex, and no trend with redshift. Although a systematic offset in the $\Delta \log L_{3000}$ measurements is expected, as our $L_{3000}$ could be contaminated by the Fe\textsc{ii} emission, this is reassuringly small (only a flux factor of 1.3). We also note that, any redshift independent offset in the UV fluxes would not be an issue, since it is balanced out in the cross-calibration between the Hubble diagram of supernovae and quasars.

Figure 4 also shows the $\Delta \log L_{3000}$ distribution as a function of redshift for the 1473 quasars in the clean SDSS–4XMM sample (1644 total sources, see Sect. 7 for details) with a measurement of $L_{3000}$. The average $\Delta \log L_{3000}$ is perfectly consistent with the initial sample, the dispersion around the mean is 0.13 dex, and again there is no trend with redshift, implying that our selection criteria have the only effect of singling out the most reliable luminosity measurements as proxy of the nuclear emission.

We finally note that, regardless of the details of the physical mechanism driving the $L_X - L_{\text{UV}}$ relation, the characteristic UV flux wavelength should be the one most closely related to the global emission of accretion disc, and might not be precisely the rest-frame 2500 Å (see Sect. 6 in Risaliti & Lusso 2019). It is even possible that nuclear emission should be combined with other AGN spectral properties (e.g. emission-line FWHM, Fe\textsc{ii} emission, this is reassuringly small (only a flux factor of 1.3). We also note that, any redshift independent offset in the UV fluxes would not be an issue, since it is balanced out in the cross-calibration between the Hubble diagram of supernovae and quasars.

Since we are still far from grasping the nature of the $L_X - L_{\text{UV}}$ relation, our photometric fluxes are seen as a more conservative representation of the broadband disc emission, capturing the “true” dependence between disc and X-ray corona in AGN in a statistical sense.

5. The rest-frame 2 keV monochromatic flux

Given the large source statistics in the SDSS–4XMM and XXL samples, a detailed X-ray spectral analysis of all the
quasars is impractical. Therefore, to compute the rest-frame 2 keV monochromatic flux \((F_X)\), we follow the same approach as described in Risaliti & Lusso (2019). For the SDSS–4XMM sample, we derived the rest-frame 2 keV fluxes and the relative (photometric) photon indices, \(\Gamma_X\) (along with their 1σ uncertainties), from the tabulated 0.5–2 keV (soft, \(F_S\)) and 2–12 keV (hard, \(F_H\)) fluxes reported in the 4XMM–DR9 serendipitous source catalogue. These band-integrated fluxes are blueshifted to the rest-frame by considering a pivot energy value of 1 keV \((E_S)\) and 3.45 keV \((E_H)\), respectively, and by assuming the same photon index used to derive the fluxes in the 4XMM catalogue (i.e. \(\Gamma_X = 1.42\)). For the soft band, the monochromatic flux at \(E_S\) is then:

\[
F_E(E_S) = F_S \frac{(2 - \Gamma_X)E_S^{1-\Gamma_X}}{(2 \text{keV})^{2-\Gamma_X} - (0.5 \text{keV})^{2-\Gamma_X}},
\]

in units of erg s\(^{-1}\) cm\(^{-2}\) keV\(^{-1}\). An equivalent expression holds for the hard band, with the obvious modifications. Flux values are corrected for Galactic absorption.

The photometric photon index is then estimated from the slope of the power law connecting the two soft and hard monochromatic fluxes at the rest-frame energies corresponding to the observed pivot points. The rest-frame photometric 2 keV flux (and its uncertainty) is interpolated (or extrapolated) based on such a power law.

To justify the employed soft and hard pivot energy values and to ensure that our photometric \(F_X\) values are accurate, we performed on the one hand several simulations, and on the other hand full X-ray spectral fitting of a number of random objects at different redshifts.

Regarding the former approach, we simulated a high-quality power-law spectrum, assuming both a typical average background and calibrations for XMM–Newton, with the same photon index assumed by the 4XMM–DR9 catalogue. We fitted the data in the soft band with a power law parametrised as \(F(E) = F(0)\frac{E}{(E/E_0)}^{-\Gamma_X}\), with \(F(0)\) and \(\Gamma_X\) as free parameters, with \(E_0\) ranging from 0.5 to 1.5 keV in steps of 0.05 keV. In each case, we derived the \(F(0)\) confidence contours. In general, the covariance between these two parameters is non-zero (i.e. the contours are elongated and tilted), so we explored which value \(\Gamma_X\) returns covariance zero between \(F(E_S)\) and \(\Gamma_X\) (i.e. the contours become 2D Gaussians). This pivot energy represents the energy value dividing the soft band in two regions having the same statistical weight. As such, this value is not located at exactly the centre of the energy band because of the dependence of the effective area on energy.

As a result of the \(F(E_S) = \Gamma_X\) zero-covariance, our photometric \(F_X\) values are independent of the specific \(\Gamma_X\) assumed in the 4XMM–DR9 catalogue. This also implies that, even if our photometric \(\Gamma_X\) deviates from the true intrinsic value, the resulting \(F_X\) will be accurate in any case. Finally, the relative error on the monochromatic flux, \(\Delta F(E)/F(E)\), at the pivot energy is the same as the one of the band flux, whereby the absolute value of \(\Delta F(E)\) at the pivot energy is the smallest possible. The same procedure is also applied to the XXL sample using their catalogued soft and hard band fluxes.

In parallel, we performed a full spectral analysis on a number of random objects. The top panels in Figs. 5 and 6 present the comparison between the inferred spectroscopic and photometric \(F_X\) and \(\Gamma_X\) values, respectively, for 30 random quasars in the SDSS–4XMM sample. We also considered for this comparison the 27 sources in the \(z \sim 3\) quasar sample (Nardini et al. 2019) that have an entry in the 4XMM–DR9 catalogue. The points are colour-coded by their number of net counts. Whilst the values of \(\Gamma_X\) display a large scatter (up to \(-0.46\) dex), our photometric \(F_X\) values are in remarkable agreement with the spectroscopic ones (with a scatter of just \(-0.15\) dex). The most obvious outlier in the bottom panel of Fig. 6 is J1425+54, a marginally detected \(z = 3\) quasar with 22 ± 13 net counts in the pn (see Table 1 in Nardini et al. 2019; for the same camera, 4XMM–DR9 gives a consistent number of counts). The observed soft flux reported in the 4XMM–DR9 catalogue for this object is \((2.0 \pm 0.5) \times 10^{-15}\) erg s\(^{-1}\) cm\(^{-2}\), whilst it is virtually undetected in the observed hard band, with a S/N of \(-0.5\) and \(F_H = (9 \pm 20) \times 10^{-16}\) erg s\(^{-1}\) cm\(^{-2}\). Nonetheless, even with a nominally large discrepancy between the spectroscopic and photometric \(\Gamma_X\) values, the \(F_X\) estimates are well within a factor of 1.2. Overall, we have a consistency within a flux factor of 1.6 for about 80% of the sample (only 12/57 quasars lie outside \(|\Delta \log F_X| = 0.2\) and, as expected, the higher the number of counts, the better the agreement, with the most deviant points having less than 100 counts.

We have also investigated whether the difference between spectroscopic and photometric fluxes (\(\Delta F_X = F_X\) spectroscopic–\(F_X\) photometric) and photon indices (\(\Delta \Gamma_X = \Gamma_X\) spectroscopic–\(\Gamma_X\) photometric) displays any trend with redshift. The bottom panels of Figs. 5 and 6 show such distributions and, despite the limited statistics, both \(\Delta F_X\) and \(\Delta \Gamma_X\) are scattered around zero with no clear trend.

The rather poor comparison between \(\Gamma_X\) spectroscopic and \(\Gamma_X\) photometric shown in Fig. 6 might cast some doubts on the reliability of the photon indices derived from the broadband (soft and hard) fluxes. However, we believe that our technique of computing photometric \(\Gamma_X\) values can be safely employed for large sample of quasars and that it provides robust results, for the following reasons: firstly, the spectroscopic and photometric X-ray fluxes are in very good agreement, meaning that our distance measures are not strongly affected by the use of photometric \(\Gamma_X\) values; and secondly, we performed a series of checks by varying the photometric \(\Gamma_X\) range employed to select the final sample, finding that our main results are not significantly modified (see Sect. 7.2 for further details). In summary, our \(\Gamma_X\) photometric may not be correct on an object-by-object basis, but they are reliable in a statistical sense for large enough samples.

6. X-ray non-detected quasars

Quasar samples that include X-ray non-detections are likely to be unbiased, but the analysis of both the \(L_{XUV} - L_X\) and the distance modulus–redshift relations is far from straightforward, since it strongly depends on the weights assumed in the fitting algorithm. In the case of flux-limited surveys, objects with an expected emission (based on the observed \(L_{XUV} - L_X\) relation) close to the flux limit will be observed only in case of positive fluctuations, and this effect is likely redshift-dependent (see Sect. 7.3). Considering only detections might thus introduce some bias in the \(L_{XUV} - L_X\) relation, and this should be more relevant to the X-rays, since the relative observed flux interval is narrower than in the UV.

8 Out of the 30 objects in the \(z \sim 3\) sample, one had no public data on 2018 December 18, and two (J0945+23 and J1159+31; see Sect. 4.3 in Nardini et al. 2019) are not detected in 4XMM–DR9.

9 Interestingly, J1425+54 would not have met the selection criterion on \(\Gamma_X\) in either case (see Sect. 7.2).

10 We note that this is the typical threshold for spectral analysis to return reasonably accurate results (see e.g. Nardini et al. 2019).
Fig. 5. **Top panel**: comparison between the spectroscopic and photometric 2 keV monochromatic fluxes for 30 quasars randomly extracted from the SDSS–4XMM sample and the quasars of the $z \approx 3$ sample (Nardini et al. 2019, marked with open star symbols). **Bottom panel**: difference between the spectroscopic and photometric 2 keV monochromatic fluxes as a function of redshift. The mean $\Delta \log F_X = \log F_{X,\text{spectro}} - \log F_{X,\text{photo}}$ value and its 1σ dispersion are shown with the red and dashed lines, respectively. Points are colour-coded by the number of net counts. The $\Delta \log F_X$ distribution is scattered around $\Delta \log F_X \sim 0$ with no clear trend with redshift.

Fig. 6. **Top panel**: comparison between the spectroscopic and photometric photon indices for 30 quasars randomly extracted from the SDSS–4XMM sample and the quasars of the $z \approx 3$ sample (Nardini et al. 2019, marked with open star symbols). **Bottom panel**: difference between the spectroscopic and photometric photon index values as a function of redshift. The mean $\Delta \Gamma_X = \Gamma_{X,\text{spectro}} - \Gamma_{X,\text{photo}}$ value and its 1σ dispersion are shown with the red and dashed lines, respectively. Points are colour-coded by the number of net counts. Although the $\Gamma_X$ distribution along the one-to-one relation is rather scattered, the $\Delta \Gamma_X$ does not seem to show a clear trend with redshift.

Lusso & Risaliti (2016) investigated the effect of the inclusion of X-ray non-detections in the study of the $L_X - L_{UV}$ relation for an optically selected sample of quasars, whose selection was very similar to the one employed in the present analysis. Their main conclusion was that there were no statistically significant variations on slope, intercept and dispersion (within their uncertainties) between X-ray detected and censored quasar samples across the different selection steps, with the slope being rather constant around 0.6. Further analysis was performed in Risaliti & Lusso (2019, see Sect. 3 of their supplementary material), where the fraction of X-ray non-detected quasars was of the order of 2% in their final cleaned sample. Such a fraction of censored data has negligible statistical weight in the fitting procedure, so their inclusion does not change the results of the statistical analysis.

Here, we have adopted a similar strategy as the one presented by Risaliti & Lusso (2019) to obtain a sample where biases are minimised even without the inclusion of non-detections, which is discussed at length in Sect. 7.3. Additionally, we explored whether any possible remaining bias in our X-ray detected quasar sample is present in the residuals of the quasar Hubble diagram (see Sect. 11.1). All these checks motivated us to analyse the Hubble diagram where non-detections are neglected.

7. Selection of a clean quasar sample

Our aim is to select a sub-sample with accurate estimates of $L_{UV}$ and $L_X$, covering a redshift range as wide as possible, by removing systematic effects and low-quality measurements. For the latter, we applied a couple of preliminary filters that ensure good measurement quality. These filters mainly involve the X-ray data, since these affected by larger uncertainties. Specifically, we considered only soft and hard flux measurements with a relative error smaller than 1 (i.e. a minimum $S/N$ of 1 on both band fluxes): $\Delta F_S / F_S < 1$ and $\Delta F_H / F_H < 1$. A similar filter is currently not applied to UV fluxes since the $S/N$ at these wavelengths is typically much higher than 1. Overall, these two filters exclude about 30% of the X-ray detections in the initial sample.
The main possible sources of contamination and systematic error are: dust reddening and host-galaxy contamination in the optical and UV, gas absorption in the X-rays, and Eddington bias associated with the flux limit of the X-ray observations. Here we briefly discuss each of these effects, and describe the filters we applied to obtain the final “best” sample for a cosmological analysis.

7.1 Dust reddening and host-galaxy contamination

To retain the quasars with minimum levels of dust reddening and host galaxy contamination, we follow a similar approach to the one presented in our previous works (Risaliti & Lusso 2015, 2019; Lusso & Risaliti 2016). We used the rest-frame photometric SEDs discussed in Sect. 3 to compute, for each object, the slope $\Gamma_1 - \Gamma_2$ of a log$(\nu) - \log(vL_*)$ power law in the rest frame 0.3–1 $\mu$m and 1450–3000 Å intervals, respectively (see Sect. 7.1). The stars represent the $\Gamma_1 - \Gamma_2$ values of the quasar SED by Richards et al. (2006) with increasing dust reddening (following the extinction law of Prevot et al. 1984), with $E(B-V)$ in the range 0–0.3. We selected all the quasars inside the blue circle (i.e. with minimum host-galaxy and dust reddening contamination).

Fig. 7. Example of the $\Gamma_1 - \Gamma_2$ distribution for the XXL quasar sample, where $\Gamma_1$ and $\Gamma_2$ are the slopes of a power law in the log$(\nu) - \log(vL_*)$ plane in the 0.3–1 $\mu$m and 1450–3000 Å intervals, respectively (see Sect. 7.1). The stars represent the $\Gamma_1 - \Gamma_2$ values of the quasar SED by Richards et al. (2006) with increasing dust reddening (following the extinction law of Prevot et al. 1984), with $E(B-V)$ in the range 0–0.3. We selected all the quasars inside the blue circle (i.e. with minimum host-galaxy and dust reddening contamination).

We selected all the quasars inside the blue circle (i.e. with minimum extinction law of Prevot et al. 1984), with $R_V = 3.1$ (as appropriate for unobscured AGN; Hopkins et al. 2004; Salvato et al. 2009), to estimate the $\Gamma_1 - \Gamma_2$ correlation as a function of extinction, parameterised by the colour excess $E(B-V)$. We obtained the red dashed line shown in Fig. 7, where the starting point corresponds to the SED of Richards et al. (2006, i.e. $\Gamma_1 = 0.82, \Gamma_2 = 0.40$) with zero extinction. The distribution of $\Gamma_1 - \Gamma_2$ towards low values along the red dashed line is indicative of possible dust reddening, whilst sources towards more negative $\Gamma_1$ values are objects with possible host-galaxy contamination. The $\Gamma_1 - \Gamma_2$ plane is also very useful to identify unusual SEDs or SEDs characterised by bad photometry, which are then excluded from the sample. We selected all the sources in the $(\Gamma_1, \Gamma_2)$ plane within a circle centred at the reference values for a standard quasar SED (see Risaliti & Lusso 2015, 2019; Lusso & Risaliti 2016 for further details), with a radius corresponding to a reddening $E(B-V) = 0.1$.

We note that our quasar selection based on photometry could still be affected by some contamination from the light of the host, especially in low redshift (z $\lesssim 0.7$) AGN, whose flux values at 2500 Å are located at the edge of the SDSS photometric coverage. Hence, any uncertainties in the estimate of the quasar UV continuum from the optical can make the 2500 Å monochromatic fluxes less reliable and possibly overestimated. Low-redshift AGN are on average less luminous in the optical and UV, with $L_{\text{tot}} \lesssim 10^{44}$ erg s$^{-1}$, thus the contrast between nuclear continuum and host-galaxy emission is smaller with respect to higher luminosity objects. Moreover, the data quality of low-redshift and low-luminosity AGN is, on average, lower. Host-galaxy contamination can be minimised through a source-by-source spectral fitting, as we did for the local AGN sample, but this procedure is rather time consuming for samples of several hundred thousands of objects. We further discuss possible issues for cosmology related to our selection in Sect. 10.

7.2 X-ray absorption

Since X-ray fluxes may contain some level of absorption, which is naturally heavier in the soft band, we included only X-ray detections with a photon index $\Gamma_X$ that falls within a range representative of unobscured quasars. For the majority of the sample, we adopted the following selection criterion, which also takes into account the uncertainties on $\Gamma_X$: $\Gamma_X - \Delta \Gamma_X \geq \Gamma_{\text{min}}$ and $\Gamma_X \leq \Gamma_{\text{max}}$. The values of $\Gamma_{\text{min}}$ and $\Gamma_{\text{max}}$ are chosen based on two considerations: the average $\Gamma_X$ within that interval should roughly correspond to $\Gamma_X \sim 2$ with a dispersion of 0.2–0.3 (consistent with e.g. Young et al. 2009a), and the $L_X - \Gamma_{\text{UV}}$ relation should not present any systematic deviation from the assumed true slope of 0.6 (within uncertainties).

We thus proceeded as follows. We evaluated the $F_X - F_{\text{UV}}$ relation in narrow redshift bins (so the effect of cosmology is negligible) for different choices of $\Gamma_{\text{min}}$ and $\Gamma_{\text{max}}$. For the SDSS–4XMM sample, we started by assuming a reasonable $\Gamma_{\text{max}}$ of 2.8 and a varying $\Gamma_{\text{min}}$ in the interval 1.4–1.9 with steps of 0.1, and converged to $\Gamma_{\text{min}} = 1.7$. We checked that a smaller value of $\Gamma_{\text{min}}$ (i.e. 1.6) would not change the results of our analysis, but we prefer to be conservative, even at the expenses of sample statistics. For the SDSS–4XMM sample, we selected only X-ray observations with a photon index satisfying the condition $\Gamma_X - \Delta \Gamma_X > 1.7$, and excluded the (few) objects with $\Gamma_X > 2.8$. The latter filter on $\Gamma_X$ is needed to avoid strong outliers (~5%) which may be due to observational issues such as incorrect background subtraction in one of the two bands. This $\Gamma_X$ interval roughly corresponds to an average $\Gamma_X \sim 2.1$–2.2 and
a dispersion of 0.3. The same selection is applied to all the other sub-samples at \( z < 4 \). For the higher redshift (\( z > 4 \)) sample, such a stringent criterion on \( \Gamma_X \) would exclude the majority of the objects, given their higher uncertainties. We thus decided to simply select all the objects with \( \Gamma_X \geq 1.7 \).

Given the observed \( \Gamma_X \) range (up to 2.8), some soft-excess (e.g. Sobolewska & Done 2007; Gliozzi & Williams 2020) contribution for low-\( z \) quasars might be still present. We have thus repeated the analysis further imposing an upper limit to the \( \Gamma_X \) range of 2.5, but, besides losing statistics, our results are not affected.

### 7.3. Eddington bias

Owing to X-ray variability, AGN with an average X-ray intensity close to the flux limit of the observation will be observed only in case of a positive fluctuation. This introduces a systematic, redshift-dependent bias towards high fluxes, known as Eddington bias, which has the effect to flatten the \( L_X - L_{UV} \) relation.

To reduce this bias, we excluded all X-ray detections below a threshold defined as a times the intrinsic dispersion of the \( L_X - L_{UV} \) relation (LR16; Risaliti & Lusso 2019), specifically:

\[
\log F_{2\text{keV,exp}} - \log F_{\text{min}} < \kappa \delta, \quad (3)
\]

where \( F_{2\text{keV,exp}} \) is the monochromatic flux at 2 keV expected from the observed rest-frame quasar flux at 2500 Å with the assumption of a true \( \gamma = 0.6 \), and it is calculated as follows:

\[
\log F_{2\text{keV,exp}} = (\gamma-1)\log(4\pi)+2(\gamma-2)\log d_l+\gamma \log F_{UV}+\beta, \quad (4)
\]

where \( d_l \) is the luminosity distance calculated for each redshift with a fixed cosmology, and the parameter \( \beta \) represents the pivot point of the non-linear relation in luminosities, \( \beta = 26.5 - 30.5y \approx 8.2^{11} \), \( F_{\text{min}} \) in Eq. (3) is obtained as detailed in Sects. 2.1 and 2.2, whilst the product \( \kappa \delta \) is a value estimated for all the sub-samples that we constructed from archives (SDSS–4XMM, SDSS–Chandra) or surveys (XXL).

In particular, we first computed the flux limit of each X-ray observation, for both the SDSS–4XMM and SDSS–Chandra samples (see Sects. 2.1 and 2.2). We minimised the Eddington bias by including only X-ray detections for which the minimum detectable flux \( F_{\text{min}} \) in that given observation is lower than the expected X-ray flux \( F_{2\text{keV,exp}} \) by a factor that is proportional to the intrinsic dispersion in the \( L_X - L_{UV} \) relation (we refer to Appendix A in LR16 and Risaliti & Lusso 2019).

On average, the minimum detectable monochromatic fluxes at 2 keV are approximately \( 4.6 \times 10^{-12} \text{erg s}^{-1} \text{cm}^{-2} \text{Hz}^{-1} \) and \( 3 \times 10^{-12} \text{erg s}^{-1} \text{cm}^{-2} \text{Hz}^{-1} \) for the SDSS–4XMM and SDSS–Chandra samples, respectively. However, we caution that these values should not be considered the “survey limiting fluxes”, since both the 4XMM and CSC2.0 catalogues are not proper flux-limited samples, but rather a collection of all X-ray observations performed over a certain period. It is thus not trivial to estimate the expected minimum flux for these catalogues. The XXL sample is, instead, a “standard” flux-limited sample, so we applied a soft-band flux threshold to the data \( (F_5 > 10^{-15} \text{erg s}^{-1} \text{cm}^{-2}) \), which corresponds to a flux limit at 2 keV of \( 5 \times 10^{-12} \text{erg s}^{-1} \text{cm}^{-2} \text{Hz}^{-1} \). We considered \( \kappa = 0.9 \) for SDSS–4XMM and XXL, whilst we used \( \kappa = 0.5 \) for the SDSS–Chandra sample. All the other sub-samples rely on pointed observations, so we did not apply any flux threshold to the data.

In principle, the effects of this bias could be further reduced if also non-detections were considered. Yet, this would not only complicate the statistical analysis, but also make the estimate of the intrinsic dispersion of the observed relations (e.g. \( L_{UV} - L_X \), Hubble diagram) much more uncertain. Moreover, we have shown that there is no significant variation in both the slope and the intercept of the \( L_{UV} - L_X \) correlation (within their uncertainties) among censored and detected samples once the Eddington bias is taken into account (see Appendix A in LR16). We therefore decided to include only detections in this work. This choice implies that we have to be very conservative in the correction for the Eddington bias, at the expense of sample statistics.

It is worth noting that our procedure to minimise the Eddington bias is slightly circular: we need the \( L_X - L_{UV} \) relation (i.e. we assumed \( y = 0.6 \)) in order to estimate the “expected” X-ray flux. Yet, our simulations show that we are able to retrieve the assumed cosmology (using different input values for \( \Omega_M \) and \( \Omega_{\Lambda} \)), when the selection criteria are applied to mock quasar samples.

### 7.4. The final cleaned sample

In summary, we applied a series of selection criteria to filter all the data that are likely contaminated by dust reddening, host-galaxy contamination, and X-ray absorption, or affected by the Eddington bias. We first selected all quasars within a circle centred at \((\Gamma_1, \Gamma_2) = (0.82, 0.4)\), that is \( E(B-V) = 0 \), and with a radius such as:

\[
\sqrt{(\Gamma_1 - 0.82)^2 + (\Gamma_2 - 0.40)^2} \lesssim 1.1, \quad (5)
\]

which corresponds to an \( E(B-V) \lesssim 0.1 \). The equation above filters out all quasar SEDs that show reddening in the UV, significant host-galaxy contamination in the near-infrared, as well as bad photometry (see Sect. 7.1). We then applied an additional cut to keep only the X-ray observations where photon indices are indicative of low levels of X-ray absorption, and to exclude the X-ray data characterised by peculiar photon indices, especially at low and moderate redshifts (see Sect. 7.2). Specifically, we required that:

\[
\begin{align*}
\Gamma_X + \delta \Gamma_X & \geq 1.7 \quad \text{and} \quad \Gamma_X \leq 2.8, \quad \text{if} \quad z < 4 \\
\Gamma_X & \geq 1.7, \quad \text{if} \quad z \geq 4.
\end{align*}
\]

To correct for the Eddington bias, we further selected all observations that satisfy Eq. (3) where the product \( \kappa \delta \) is 0.9 for the SDSS–4XMM and XXL sub-samples, and 0.5 for SDSS–Chandra. Pointed observations are available for the local, \( z \approx 3 \), and high-redshift samples (see Sect. 7.3). For any quasar, all the multiple X-ray observations that survive the filters above are finally averaged to minimise the effects of X-ray variability (e.g. LR16, see also Lusso 2019).

The final cleaned sample is composed by 2421 quasars spanning a redshift interval \( 0.009 \leq z \leq 7.52 \), with a mean (median) redshift of 1.442 (1.295). Table 1 summarises the statistics of each sub-sample, whilst a more detailed summary of the various sub-samples after a given selection is provided in Table A1. The main UV and X-ray properties of the final sample are presented in Table 3.

### 8. Analysis of the \( F_X - F_{UV} \) relation with redshift

Before building the Hubble diagram, we need to check whether the \( F_X - F_{UV} \) relation for the clean quasar sample shows any
Table 3. Properties of the final quasar sample.

| Name       | RA (deg) | Dec (deg) | z     | log $F_{\text{UV}}$ | log $F_X$ | Group | $\Gamma_X$ | DM           |
|------------|----------|-----------|-------|---------------------|-----------|--------|------------|--------------|
| 030341.04–002321.9 | 45.92103 | −0.58942 | 3.235 | −27.00 ± 0.04       | −31.38 ± 0.04 | 1      | 1.87 ± 0.12 | 46.42 ± 0.31 |
| 030449.85–000813.4 | 46.20775 | −0.13708 | 3.296 | −26.98 ± 0.01       | −31.24 ± 0.03 | 1      | 1.99 ± 0.29 | 45.62 ± 0.20 |
| 090508.88+305757.3 | 136.28702 | 30.96593 | 3.034 | −26.97 ± 0.01       | −31.14 ± 0.02 | 1      | 2.12 ± 0.09 | 44.90 ± 0.17 |

Notes. This table is presented in its entirety at the CDS; a portion is shown here for guidance. Fluxes are in units of log(erg s$^{-1}$ cm$^{-2}$ Hz$^{-1}$). The Group column flags the different sub-samples: 1 = XMM–Newton $z \geq 3$, 2 = new XMM–Newton $z \approx 4$ quasar, 3 = High-$z$ sample, 4 = XXL, 5 = SDSS – 4XMM, 6 = SDSS – Chandra, 7 = local AGN. The column DM reports the distance moduli (with uncertainties) to reproduce Fig. 9.

Fig. 8. Redshift evolution of the slope $\gamma$, the intercept $\beta$ and the dispersion $\delta$ of the $F_X$–$F_{\text{UV}}$ relation. To perform the regression fit, X-ray and UV fluxes have been normalised to $10^{-28}$ and $10^{-30}$ erg s$^{-1}$ cm$^{-2}$ Hz$^{-1}$, respectively. The data points in each panel are averages in narrow redshift bins ($\Delta z = 0.06$–0.45). Error bars represent the $1\sigma$ uncertainty on the mean in each bin. The grey solid and dashed lines are the means and $1\sigma$ uncertainties, respectively, on the slope ($\gamma$) and on the dispersion ($\delta$). The red line marks $\gamma = 0.6$.

The trend with redshift. We thus divided the sample in narrow redshift bins, with a variable step $\Delta z = 0.06$–0.45 within the redshift range 0.45–4 to have enough statistics. The redshift step is chosen to ensure that the dispersion in distances over each interval is smaller than the one of the relation in luminosities. In this way, we can consider fluxes as proxies of luminosities.

The best-fit parameters (slope, intercept and dispersion) of the $F_X$–$F_{\text{UV}}$ relation and their uncertainties are shown in Fig. 8, whilst all the fits of the $F_X$–$F_{\text{UV}}$ relation in the chosen redshift bins are presented in Fig. B.1. They are obtained through the Python package EMCEE (Foreman-Mackey et al. 2013), which is a pure-Python implementation of Goodman & Weare’s affine invariant Markov chain Monte Carlo (MCMC) ensemble sampler. To perform the regression fit, X-ray and UV fluxes were normalised to $10^{-28}$ and $10^{-30}$ erg s$^{-1}$ cm$^{-2}$ Hz$^{-1}$, respectively. On average, the $F_X$–$F_{\text{UV}}$ slope does not show any clear trend with redshift within the analysed interval. Conversely, the trend of the intercept $\beta$ of the normalised $F_X$–$F_{\text{UV}}$ relation observed in the middle panel of Fig. 8 just reflects the overall shape of the quasar Hubble diagram (see Sect. 9). We note that, the trend of $\beta$ with redshift, is not exactly the same as the one in Fig. 9 since such a parameter does not have a simple direct proportional dependence on the distance modulus (Eq. (7)) because of the different dependence between slope and normalisation in each redshift bin.

The sample statistics is so sparse at a redshift higher than 4 that we cannot provide a meaningful fit of the relation. Nonetheless, we have checked that the data points at $z > 4$ do not show any trend with redshift in the residuals of the Hubble diagram (see Sect. 11). In fact, these data points are extremely useful to set the shape of the Hubble diagram, thus providing better constraints on the measurements of the expansion rate of the Universe. We thus confirm that the slope of the X-ray to UV relation shows no redshift evolution up to $z \sim 4$, in agreement with our previous works (e.g. Risaliti & Lusso 2015, 2019; Lusso & Risaliti 2016, 2017; Lusso et al. 2019; see also Salvestrini et al. 2019 for a high redshift analysis).

9. The quasar Hubble diagram

To fit the Hubble diagram we first need to derive the distance modulus for each object. We start by computing the luminosity distance (e.g. see Risaliti & Lusso 2015, 2019) as:

$$\log d_L = \frac{\left[\log F_X - \beta - \gamma(\log F_{\text{UV}} + 27.5)\right]}{2(\gamma - 1)} +$$

$$- \frac{1}{2} \log(4\pi) + 28.5,$$

where $F_X$ and $F_{\text{UV}}$ are the flux densities (in erg s$^{-1}$ cm$^{-2}$ Hz$^{-1}$), $F_{\text{UV}}$ is normalised to the (logarithmic) value of 27.5 in the equation above, whilst $d_L$ is in units of cm and is normalised to 28.5 (in logarithm). The slope of the $F_X$–$F_{\text{UV}}$ relation, $\gamma$, is a free parameter, and so is the intercept $\beta$. The distance modulus, $dM$, is thus:

$$dM = 5 \log d_L - 5 \log(10 \text{pc}),$$

and the uncertainty on $dM$, $dDM$, is:

$$dDM = \frac{5}{2(\gamma - 1)} \left(\frac{d \log F_X}{\log(4\pi)} + \frac{d \log F_{\text{UV}}}{\log(4\pi)} + \frac{(d\beta)^2}{(\gamma - 1)} \right)^{1/2},$$

where $d \log F_X$ and $d \log F_{\text{UV}}$ are the logarithmic uncertainties on $F_X$ and $F_{\text{UV}}$, respectively. Equation (9) assumes that all the parameters are independent, and takes into account also the uncertainties on $\beta$ and $\gamma$. The fitted likelihood function, $L_F$, is

---

12 The intercept $\beta$ of the $L_X$–$L_{\text{UV}}$ relation is related to the one of the $F_X$–$F_{\text{UV}}$, $\beta$ (see Sect. 8), as $\beta(z) = 2(\gamma - 1) \log d_L(z) + (\gamma - 1) \log 4\pi + \beta_z$.
then defined as:
\[
\ln LF = -\frac{1}{2} \sum_{i=1}^{N} \left( \frac{(y_i - \psi_i)^2}{s_i^2} - \ln s_i^2 \right)
\]  

where \(N\) is the number of sources, \(s_i^2 = d_i^2 + \gamma_i^2 d_i^2 + \exp(2 \ln \delta)\) takes into account the uncertainties on both the \(x_i\) (\(\log F_{\nu,i}\)) and \(y_i\) (\(\log F_X\)) parameters of the fitted relation, whilst \(\delta\) represents its intrinsic dispersion. The variable \(\psi\) is the modelled X-ray monochromatic flux, defined as:
\[
\psi = \log F_{X,\text{mod}} = \beta + \gamma (\log F_{\nu,\text{mod}} + 27.5) + 2(\gamma - 1)(\log d_{L,\text{mod}} - 28.5),
\]

and is dependent upon the data, the redshift and the cosmological model assumed for the distances (e.g. \(\Lambda\)CDM, \(w\)CDM or a polynomial function). We fitted the data with a luminosity distance described by a fifth-grade polynomial of \(\log(1+z)\), where the cosmographic function is:

\[
d(\zeta)_{L,\text{mod}} = k \ln(10) \frac{c}{H_0} \left[ \log(1+z) + a_2 \log^2(1+z) + a_1 \log^3(1+z) + a_0 \log^5(1+z) \right] + O(\log^6(1+z)),
\]

where \(k, a_0, a_1, a_2, a_3\), and \(a_4\) are free parameters.

For any analysis that involves a detailed test of cosmological models, we should cross-calibrate quasar distances making use of the distance ladder through supernovae Ia. In fact, the DM values of quasars are not absolute, thus a cross-calibration parameter \(k\) is needed. The parameter \(k\) should be fit separately for supernovae Ia and quasars (i.e. \(k\) is a rigid shift of the quasar Hubble diagram to match the one of supernovae).

Whilst in our previous works we kept \(k\) fixed, in this analysis we have marginalised over the slope \(\gamma\) of the \(L_X - \log F_{\nu}\) relation. The latter approach is preferred to check whether any degeneracy of the slope with the other parameters is present, and whether the statistical significance of the deviation from the \(\Lambda\)CDM model can be affected by the assumption of a \(\gamma\) value that slightly deviates from the true one. The marginalisation on \(\gamma\) is a more conservative procedure, hence it might reduce the significance of the deviation with respect to the same MCMC analysis with \(\gamma\) fixed. Therefore, if a statistical deviation persists even allowing for a variable \(\gamma\), its significance should be considered as an indicative lower limit with respect to the case where \(\gamma\) is fixed.

We finally note that the Hubble constant \(H_0\) in Eq. (12) is degenerate with the \(k\) parameter, so it can assume any arbitrary value and was fixed to \(H_0 = 70\) \(\text{km s}^{-1} \text{Mpc}^{-1}\) (see also Lusso et al. 2019).

Figure 9 shows the Hubble diagram for the clean quasar sample, combined with the most updated compilation of supernovae Ia from the Pantheon survey (Sconlic et al. 2018). The best MCMC cosmographic fit is shown with the red line, whilst black points are the averages (along with their uncertainties) of the distance moduli in narrow (logarithmic) redshift intervals, plotted for clarity purposes only. The residuals are displayed in the middle panel with the same symbols, and do not reveal any apparent trend with redshift. The MCMC fit assumes uniform priors on the parameters. More details on our cosmographic technique will be provided in a companion publication (Bargiacchi et al., in prep.).

We confirm that, whilst the Hubble diagram of quasars is well reproduced by a standard flat \(\Lambda\)CDM model (with \(\Omega_M = 0.3\)) up to \(z \sim 1.5\), as shown in the top panel of Fig. 9, a statistically significant deviation emerges at higher redshifts, in agreement with our previous works (e.g. Risaliti & Lusso 2015, 2019; Lusso et al. 2019) and other works on the same topic (e.g. Di Valentino et al. 2020).

The detailed discussion of the cosmological implications of this deviation and its statistical significance is not the main aim of this analysis. Here we want to focus on the study of possible systematic effects that could drive this deviation instead.

10. Cosmological fits of the Hubble diagram

In this section we want to test our quasar sample and our method by fitting a “physical” cosmological model. Our aim is not to fully explore the consequences of our new Hubble diagram for
the determination of cosmological parameters, or for the tests of different cosmological models, which will be presented in subsequent papers. Here we only intend to verify how different choices regarding the fitting method and the quasar subsample affect the final results. We choose to perform these tests with a flat $w_0\Lambda_0$CDM model, which is the simplest and most commonly used extension of the standard $\Lambda$CDM model, where the parameter $w$ of the equation of state of dark energy is assumed to vary with redshift according to the parametrisation \[ w(z) = w_0 + w_a \times (1 - a), \] where $a = 1/(1 + z)$. Based on the analysis presented in the previous sections, three points deserve further consideration regarding cosmological fits.

Firstly, we can use the full quasar sample or add a filter of $z > 0.7$ for the sources with photometric determination of the UV flux. As discussed in Sect. 7.1, the possible uncertainties in the extrapolation from the optical of the quasar UV continuum at low redshift, where the host-galaxy contamination can be important, make the 2500 Å monochromatic luminosities less reliable at $z < 0.7$. In particular, this effect is likely to be more severe at lower fluxes (or luminosities), where the data quality is also lower. If the continuum slope at UV wavelengths becomes steeper than in the optical, the actual 2500 Å flux would become steeper than in the optical, the actual 2500 Å flux would be underestimated and this could explain the higher average values of the slope of the $L_X - L_{UV}$ relation at $z < 0.7$ (see Fig. 8). This point deserves further investigation, which is deferred to a subsequent paper. To define the optical sample for cosmological applications in a conservative way, we thus prefer to cut the quasar sample at $z > 0.7$, with the exception of the local sources discussed in Sect. 2.7, whose 2500 Å flux is determined from the UV spectra without extrapolations. The results of the fit of this best quasar sample with the flat $w_0\Lambda_0$CDM model are shown in Fig. 10. We can see that considering our data significantly reduces the $w_0 - w_a$ parametric space with respect to the CMB analysis only (Planck Collaboration VI 2020)\(^{13}\), still being compatible with the latter data at $1\sigma$. At the same time, the $\Lambda$CDM model (recovered for values $w_0 = -1$ and $w_a = 0$) is in tension with our data at more than $3\sigma$, in agreement with RL19. We finally note that the role of quasars at $z < 1.3$ is mainly to set the absolute calibration with supernovae in cosmological fits, with only a small contribution to the determination of the values and uncertainties of the cosmological parameters, given the much higher statistical weight of the supernovae. Removing quasars at $z < 0.7$ should not affect the final results significantly, and the number of quasars at redshifts overlapping with supernovae remains high enough for a precise calibration. In order to test these expectations, we repeated the fit with the whole sample, obtaining the results shown in the first panel of Fig. C.1, where the contours are nearly indistinguishable from those of Fig. 10.

Secondly, to fit the Hubble diagram, we can either adopt a fixed value of the slope of the relation (with its uncertainty), $\gamma = 0.59 \pm 0.06$, as determined in Sect. 8, or marginalise on $\gamma$ as a free parameter, as discussed in Sect. 9. In general, the latter choice is more conservative and should be preferred. This is what we did for our reference fit, and also for the cosmographic fit used to analyse the residuals. Yet, it is worth noting that, in case of a mismatch in the shape of the Hubble diagrams of quasars and supernovae in the common redshift interval, leaving $\gamma$ as a free parameter allows us to partly alleviate this problem by slightly “bending” the $L_X - L_{UV}$ relation in order to obtain a better agreement. In our case, the fit of the best sample gives $\gamma = 0.600 \pm 0.015$, consistent within $1.5\sigma$ with the value found from the fits of the relation in narrow redshift bins. In order to check the possible effects of this choice, we repeated the fit with the $w_0\Lambda_0$CDM model, again obtaining a contour plot totally consistent with the reference case (second panel in Fig. C.1).

Third and last consideration, the residuals in the Hubble diagram show a moderate, but statistically significant, redshift dependence on the X-ray slope, $\Gamma_X$. It is important to understand whether this introduces a bias in the fits of the Hubble diagram. We checked this possibility by splitting the sample in two parts, with $\Gamma_X < 2.2$ and $\Gamma_X > 2.2$, respectively, and repeating the fit with the $w_0\Lambda_0$CDM model. The results are shown in the last two panels of Fig. C.1. These contours are slightly larger than in the previous cases, as expected given the lower statistics, but no systematic trend is observed.

We conclude that, whilst the possible dependence on $\Gamma_X$ deserves further analysis in order to understand its physical and observational origin and to reduce the dispersion of the $L_X - L_{UV}$ relation (Signorini et al., in prep.), no systematic effect related to $\Gamma_X$ is introduced in the Hubble diagram of quasars.

\footnote{Baseline $\Lambda$CDM chains with baseline likelihoods: https://wiki.cosmos.esa.int/planck-legacy-archive/index.php/Cosmological_Parameters}

11. Study of systematics in the Hubble diagram

Since the main aim of our analysis is to check whether any systematic is present in the residuals of the quasar Hubble diagram, at this stage we avoid the inclusion of supernovae Ia. As noted above, when only quasars are involved the DM values should not be considered as proper absolute distances. In this Section we thus present an in-depth investigation of possible systematics in the residuals of the quasar Hubble diagram, unaccounted for in the selection of the sample. In particular, we explored whether our procedure to correct for the Eddington bias (Sect. 7.3), to neglect quasars with possible gas absorptions (Sect. 7.2), and to select blue quasars based on their SED shape, where dust absorption and host-galaxy contamination are minimised (Sect. 7.1), introduces spurious trends in the Hubble diagram residuals as a function of redshift and for different intervals of the relevant parameters. For each variable, we divided the sample between the sources that fall below and above the average

**Fig. 10.** Results from a fit of a $w_0\Lambda_0$CDM model to the combined Hubble diagram of supernovae and the ‘best’ quasar sample, that is removing objects at $z < 0.7$ with a photometric determination of the UV flux (see the text for details). The green contours refer to the CMB results from Planck (Planck Collaboration VI 2020). The orange ($3\sigma$), red ($2\sigma$), and brown ($1\sigma$) contours are obtained by adding the constraints from the Hubble diagram of supernovae and quasars.
value of the variable itself, and examined each subset separately as any hidden dependence should lead to a systematic difference between the two.

11.1. Residuals as a function of the Eddington bias

To verify whether our adopted technique to correct for the Eddington bias, based on the assumption that the true slope of the $L_X - L_{UV}$ is $y = 0.6$, leaves some hidden trends in the residuals of the Hubble diagram as a function of redshift, we defined an Eddington bias parameter, $\Theta$, as the difference between the expected X-ray monochromatic flux at 2 keV and the sum of the flux sensitivity value at 2 keV ($F_{\text{min}}$) and the product $k\delta$ (see Eq. (3)):

$$\Theta = \log F_{2\text{keV,exp}} - (\log F_{\text{min}} + k\delta).$$

(13)

Given the fact that the source has survived the Eddington bias filter (see Eq. (3), where $k = 0.5$ for SDSS–Chandra and 0.9 for XXL and SDSS–4XMM, respectively), $\Theta$ is always positive: the higher its value, the lower the bias due to the flux limit of the specific X-ray observation. The other samples benefit from pointed observations, for which the bias due to the X-ray flux sensitivity is fully negligible.

We then defined two subsets, above (1069 objects) and below (1289) the mean of this distribution, $\langle \Theta \rangle = 1.17\pm0.01$ (with a 1$\sigma$ dispersion of 0.33), and plotted the residuals of the Hubble diagram as a function of redshift in Fig. 11. The three quasar samples span a rather large redshift interval, showing no clear trend of the residuals with a distance and a similar dispersion around zero ($\approx 1$), with an average residual value of about $-0.07$ (0.04) for $\Theta > 1.2$ ($\Theta < 1.2$). These results imply that our selection of X-ray observations discussed in Sect. 7.3 does not introduce any systematic trends, at least up to $z \approx 4$.

11.2. Residuals as a function of the photon index

Gas absorption in the X-ray band and strong outliers with extremely steep $\Gamma_X$ are minimised by neglecting all quasars outside a given interval of photon index values ($1.7 \leq \Gamma_X \leq 2.8$, see Sect. 7.2). In Fig. 12 we present the distribution of the $\Gamma_X$ values for the clean quasar sample. The mean (median) value for the sample is $\langle \Gamma_X \rangle = 2.202 \pm 0.005$ ($\langle \Gamma_X \rangle = 2.173^{+0.003}_{-0.006}$), with a dispersion of about 0.23 (0.30). Statistical errors are quoted.

The average $\Gamma_X$ is biased towards slightly steeper values with respect to the more typical $\Gamma_X \sim 1.9–2$ (e.g. $\Gamma_X = 1.99 \pm 0.01$ with a dispersion of 0.3, Scott et al. 2011; see also Young et al. 2009b; Mateos et al. 2010), which is likely due to our conservative cut at $\Gamma_X = 1.7$ and to the presence of a tail of sources with power laws softer than $\Gamma_X = 2.6$. The residuals of the Hubble diagram do not show a significant trend as a function of redshift, the sample is split in two subsets with $\Gamma_X$ higher and lower than 2.2 (Fig. 13). Nonetheless, the marginalised distribution of residuals shown in Fig. 14 presents an offset with respect to zero, with an average value for the two sub-samples of $-0.32 \pm 0.03$ and $0.24 \pm 0.03$ for $\Gamma_X$ higher and lower than 2.2, respectively. The dispersion around the average values is 0.82 (1.05) for $\Gamma_X > 2.2$ ($\Gamma_X < 2.2$).

As we pointed out in Sect. 7.2, our photometric $\Gamma_X$ values may not be always accurate for individual objects, but they are reliable in a statistical sense over a large sample of quasars. Consequently, the main drawback of not using the spectroscopic values is likely to increase the dispersion, rather than to introduce a strong systematic with redshift that may affect the cosmological analysis. Moreover, even considering the presence of a small redshift trend in Fig. 13 for $z < 2$, this is counterbalanced by the use of supernovae Ia in the same redshift interval. In fact, the statistical significance of the deviation of the quasar Hubble diagram from the standard $\Lambda$CDM reported in our previous works starts to be important for $z > 2$, where all the residuals do not present any obvious systematic with redshift.

11.3. Residuals as a function of SED colours

As a further check, we have explored whether the SED colours produce some systematic trends with redshift in the residuals of the Hubble diagram. We have selected two sub-samples in different regions of the ($\Gamma_1, \Gamma_2$) plane (see Fig. 7). The first one considers all the quasars within the circle having a radius of 0.45 (i.e. $\sqrt{(\Gamma_1 - 0.82)^2 + (\Gamma_2 - 0.40)^2} \leq 0.45$; 1271 quasars), which corresponds to a colour excess $E(B - V)$ of 0.04 (using an SMC-like reddening law). The second is the contiguous annulus with outer radius of 1.1 (1087 quasars) corresponding to an $E(B - V) \approx 0.1$. As the majority of the quasars in the final sample are drawn from...
Fig. 12. Distribution of the photon index $\Gamma_X$ for the clean final sample. The mean and the statistical error on the mean are also quoted. The red (blue) solid line represents the mean (median) of $\Gamma_X$, with its 1$\sigma$ dispersion around the average marked with the dashed (dot-dashed, 16% and 84%) lines.

Fig. 13. Distribution of the Hubble diagram residuals (middle panel of Fig. 9) as a function of redshift for the quasars with $\Gamma_X$ lower (higher) than the average (i.e. $\langle \Gamma_X \rangle = 2.2$, see Fig. 12). The red and blue solid lines represent the mean of the residuals in each $\Gamma_X$ distribution, respectively (averages are reported on top).

Fig. 14. Distribution of the Hubble diagram residuals (middle panel of Fig. 9) for the quasars with $\Gamma_X$ higher (blue) and lower (red) than the average (i.e. $\langle \Gamma_X \rangle = 2.2$, see Fig. 12). The blue and red solid lines represent the mean of the residuals in each $\Gamma_X$ distribution, respectively (averages are reported on top).

the SDSS–DR14 catalogue, the sub-sample with a colour excess lower than 0.04 should represents the bluest objects, whilst the sample in the outer annulus with 0.04 $\leq E(B-V) \leq 0.1$ could show some pattern in the residuals in case of a colour-related! systematic with redshift.

Figure 15 presents the distribution of the residuals as a function of redshift for the two sub-samples defined as above for the clean SDSS–4XMM, XXL and SDSS–Chandra samples, since we can construct homogeneous SEDs (and thus retrieve $\Gamma_1$ and $\Gamma_2$ slopes) for them all using the SDSS photometry from the DR14 quasar catalogue. It is clear that there is no obvious trend of the residuals with redshift in either sub-sample, implying that our colour selection does not introduce any redshift-dependent bias in the Hubble diagram. Moreover, in both cases there is no difference in the dispersion of the residuals around zero (=1), with average values of about $-0.04$ for $E(B-V) \leq 0.04$ and 0.03 for $0.04 \leq E(B-V) \leq 0.1$. This further proves that even the inclusion of sources with possible (yet modest) contamination from dust and/or host galaxy does not lead to any systematics with redshift up to $z \approx 4$.

11.4. Determining whether dust and gas absorption drives the deviation in the quasar Hubble diagram

Among the possible residual (and redshift-dependent) observational systematics in the Hubble diagram, we must also consider the presence of an additional contribution of dust reddening in the UV band. As we move to higher redshifts, the rest-frame optical and UV spectra shift to higher (shorter) frequencies (wavelengths), where the dust absorption cross-section is higher. This might lead to an underestimate of $F_{\text{UV}}$, which would imply an intrinsically larger value of the luminosity distance (and thus DM) than the one we measured.

We thus evaluated the amount of dust extinction required to make the cosmographic fit shown in Fig. 9 coincide with a standard flat $\Lambda$CDM model with $\Omega_M = 0.3$ (i.e. the black dashed line in Fig. 9). We defined five redshift intervals with increasing values of $E(B-V)$, specifically, $E(B-V) = 0, 0.1, 0.15, 0.2, \text{ and } 0.25$ at $z = 0–1$, $z = 1–2$, $z = 2–3$, $z = 3–4$, and $z > 4$. We then
Fig. 15. Distribution of the Hubble diagram residuals (middle panel of Fig. 9) as a function of redshift for the quasars with $\sqrt{(r_1 - 0.82)^2 + (r_2 - 0.40)^2} \leq 0.45$, that is $E(B - V) \leq 0.04$ (top panel, the bluest quasars in the clean sample) and $0.45 \leq \sqrt{(r_1 - 0.82)^2 + (r_2 - 0.40)^2} \leq 1.1$, that is $0.04 \leq E(B - V) \leq 0.1$ (bottom panel) in the clean SDSS–4XMM, XXL and SDSS–Chandra samples. The black and red points represent the mean and median of the residuals in narrow redshift intervals, respectively. Symbol keys as in Fig. 1.

Fig. 16. Analysis of the possible contribution of dust at 2500 Å. Top: distribution of the difference of the distance moduli ($\Delta DM$) as a function of redshift. The parameter $\Delta DM$ is defined as $DM_{\text{red}} - DM$, where $DM$ is computed from the quasar data as described in Sect. 9. The $DM_{\text{red}}$ values are estimated by correcting the observed $F_{\text{UV}}$ for an additional intrinsic (and redshift-dependent) extinction. $\tau$ is the optical depth at the rest-frame 2500 Å that would be needed to ascribe the discrepancy to an underestimated dust extinction in the UV flux. The black points represent the average (and 1σ uncertainties) of $DM$ in the same narrow redshift intervals used in Figs. 8 and 9. Middle: deviation in $\sigma$ of the average points with respect to $\Lambda$CDM. Bottom: ratio between the cosmographic fit obtained from the data, where $F_{\text{UV}}$ is corrected for additional extinction, and the $\Lambda$CDM. The cosmographic fit perfectly matches $\Lambda$CDM for the chosen values of dust reddening (see Sect. 11.4).

assumed the standard extinction law by Prevot et al. (1984) with $R_V = 3.1$ and corrected $F_{\text{UV}}$ by the amount dictated by the chosen $E(B - V)$ in each redshift interval, and we finally fitted the “reddening corrected” distance modulus ($DM_{\text{red}}$)–redshift relation.

Figure 16 shows the resulting distribution of the difference of the distance moduli, $\Delta DM$ (red points) between the $DM_{\text{red}}$ values computed above and the observed DM in Fig. 9 (see Sect. 9), as a function of redshift. The cosmographic fit obtained from the $DM_{\text{red}} - z$ relation perfectly matches the $\Lambda$CDM curve (bottom panel). We also report the values of optical depth, $\tau$, at the rest-frame 2500 Å that would be needed to entirely ascribe the discrepancy to underestimated dust extinction on the 2500 Å quasar fluxes.

We note that the negative $\Delta DM$ values at $z < 1$ (where no dust extinction correction is applied) are simply caused by the different overall cosmographic fit obtained for the $DM_{\text{red}} - z$ relation with respect to the observed one. Therefore, the $\Delta DM$ values in this redshift range should not be taken at face value.

At 2500 Å, it is $k(\lambda) \approx 7$, which corresponds to an increase in flux by a factor of $\approx 2.6 - 5$ for $E(B - V) = 0.15 - 0.25$ at $z \gtrsim 2$. The bolometric luminosities of the $z > 4$ quasars in the clean sample are in the range $10^{47.3 - 49}$ erg s$^{-1}$, assuming a bolometric correction of 2.75 at 2500 Å (Krawczyk et al. 2013). If the extra correction were effectively required, this would imply an intrinsic bolometric luminosity for these sources of the order of $10^{47.3 - 49}$ erg s$^{-1}$. For example, the highest redshift quasar in the sample, ULAS J134208.10+092838.61, would have a bolometric luminosity by at least a factor of 4 higher than reported in the literature (i.e., $1.6 \times 10^{47}$ erg s$^{-1}$, Bañados et al. 2018). This would also imply a larger black-hole mass, thus a heavier seed, which needs to be interpreted within current models of black-hole growth (e.g., Latif et al. 2013; Dijkstra et al. 2014; Pacucci et al. 2015).

For the spectroscopic XMM–Newton $z \approx 3$ sample, the $F_{\text{UV}}$ values would increase by a factor of at least 3.5, thus shifting $L_{\text{bol}}$ into the range $10^{48 - 48.8}$ erg s$^{-1}$. We note that the average stack of their SDSS spectra does not suggest any significant levels of dust absorption when compared to other composites obtained from bright, blue quasar spectra (see Fig. 2 in Nardini et al. 2019). As a result, the presence of any additional dust component does not seem to be justified.

For completeness, we have also employed a different reddening law (i.e., Fitzpatrick 1999), and modified both the $E(B - V)$ values and the redshift intervals (always verifying that the
cosmographic fit of the modified Hubble diagram is consistent with the standard ΛCDM), finding equivalent results.

In conclusion, the intrinsic UV fluxes required to offset the observed tension between the cosmographic fit of the Hubble diagram and the ΛCDM through redshift-dependent dust absorption should be, on average, brighter by 0.4–0.8 mag. This would imply a steeper extrapolated slope than observed at longer wavelengths, which would be hard to reconcile with a standard α-disc model (Shakura & Sunyaev 1976).

We note that concurrent, and redshift-dependent, absorption is found at X-ray energies. The observed 0.5–2 keV energy band moves towards higher rest-frame energies as the redshift increases, where the effect of any gas absorption should become progressively negligible. Spectra with flat photon indices (ΓX < 1.7) are removed from the sample, so we expect a minimal degree of residual gas absorption in the X-rays. In the high-redshift (z > 6) quasars for which a spectral analysis was carried out (see e.g. Vito et al. 2019), the level of intrinsic gas absorption is always suggested to be rather low (of the order of 1021 cm−2 or less). Assuming the presence of a local column density with NHI = 1023 cm−2 at z = 6, the exact correction to the inferred rest-frame 2 keV flux density depends on the details (flux, observed slope, data quality) of the single observation, but it is expected to be within a factor of 2, thus comparable to the typical measurement uncertainty. Such high NHI values are not required in any of the X-ray spectral fits for the high-redshift quasars that belong to the clean sample. By shifting the data to higher DM values, this correction would have the effect of increasing the departure from the ΛCDM (i.e. opposite to the one for the dust absorption in the UV).

11.5. On the possible dependence of the Lx − LUV relation on black-hole mass and accretion rate

One can argue that the observed Lx − LUV relation might be a secondary manifestation of some other, more fundamental, relations, involving, for instance, a possible dependence on black-hole mass and accretion rate (usually parametrised by the Eddington ratio, λedd, defined as the ratio of the bolometric and the Eddington luminosity). Yet, when estimated from single-epoch spectroscopy, both parameters are derived quantities, that is a combination of continuum luminosity and emission-line FWHM. Therefore, the presence or absence of a correlation between the residuals of the Hubble diagram and MBH and (or) the Eddington ratio could be misleading and potentially hide systematics.

Since the Lx − LUV relation already has a dependence on the nuclear luminosity (which we have discussed in detail in Sect. 4), we can explore possible additional correlations with the FWHM of a given emission line. This issue was already investigated in Lusso & Risaliti (2017), who found that not only is the dependence of the Lx − LUV relation on the FWHM (of Mg II) statistically significant, but also that such a dependence has the effect of further reducing the dispersion of the Lx − LUV correlation. However, building a statistically significant quasar sample that both covers a wide redshift range (as the one presented here) and relies on a measurement of FWHM from a single line (to ensure homogeneity in the estimate of the line parameters, see Sect. 2.1 in Lusso & Risaliti 2017) is impractical, as it would require a series of dedicated spectroscopic near-infrared campaigns to probe the same line (e.g. Mg II) at high redshifts (z > 2.5). The statistical relevance of the supplementary dependence of the Lx − LUV relation on the FWHM, as well as on other parameters inferred from spectroscopy (e.g. continuum slope, iron content, line strength) will be further investigated in a forthcoming publication, entirely dedicated to the physics of the relation.

Whilst here we prefer to consider the simple Lx − LUV relation for cosmological purposes, as the present analysis aims at verifying that our selection criteria do not introduce systematics in the quasar Hubble diagram, we can still check whether the sources in the clean sample with a measurement of the FWHM from Mg II (which provides the widest possible redshift coverage in BOSS) show any systematic trend and (or) a reduced scatter in the residuals. We thus cross-match the clean sample with the catalogue of spectral quasar properties of Rakshit et al. (2020), and selected the quasars with a broad (FWHM > 2000 km s−1) component of the Mg II emission line, finding 1858 quasars (77% of the clean sample).

The distribution of the residuals as function redshift for this sample is shown in Fig. 17. The amplitude of the scatter of the residuals is similar to the other cases discussed in Sects. 11.1–11.3. Moreover, there is no trend with redshift up to z ≃ 2.5. These results further prove that our selection criteria are already effective in selecting an optimal sample for cosmology, without introducing strong systematics.

12. Discussion and conclusions

Our group has presented a new technique that makes use of the observed non-linear relation between the 2500 Å and the 2 keV emission in quasars to provide an independent measurement of their distances, thus turning quasars into standardizable candles. Our method extends the distance modulus–redshift relation already eX−LUV of a given emission line. This issue was already investigated in Lusso & Risaliti (2017), who found that not only is the dependence of the Lx − LUV relation on the FWHM (of Mg II) statistically significant, but also that such a dependence has the effect of further reducing the dispersion of the Lx − LUV correlation. However, building a statistically significant quasar sample that both covers a wide redshift range (as the one presented here) and relies on a measurement of FWHM from a single line (to ensure homogeneity in the estimate of the line parameters, see Sect. 2.1 in Lusso & Risaliti 2017) is impractical, as it would require a series of dedicated spectroscopic near-infrared campaigns to probe the same line (e.g. Mg II) at high redshifts (z > 2.5). The statistical relevance of the supplementary dependence of the Lx − LUV relation on the FWHM, as well as on other parameters inferred from spectroscopy (e.g. continuum

![Fig. 17. Distribution of the Hubble diagram residuals (middle panel of Fig. 9) as a function of redshift for the quasars with a broad (FWHM > 2000 km s−1) component of the Mg II emission line in the clean sample. The black and red points represent the mean and median of the residuals in narrow redshift intervals, respectively. Symbol keys as in Fig. 1.](image-url)
possible, the dispersion in the $L_X - L_{UV}$ relation drops to $\approx 0.2$ dex (Lusso & Risaliti 2016, 2017).

We have previously demonstrated that the distance modulus–redshift relation of quasars at $z < 1.4$ is in agreement with that of supernovae Ia and with the concordance $\Lambda$CDM model (Risaliti & Lusso 2015, 2019; Lusso et al. 2019), yet a deviation from the $\Lambda$CDM emerges at higher redshift, with a statistical significance of about 4$\sigma$. If we interpret the latter result by considering an evolution of the dark energy equation of state in the form $w(z) = w_0 + w_a \times z/(1 + z)$, the data suggest that the dark energy density is increasing with time (Risaliti & Lusso 2019; Lusso et al. 2019). However, our technique may still have some limitations, and we need to verify that the observed deviation from the $\Lambda$CDM at a redshift $> 2$ is not driven by either systematics in the quasar sample selection or the cosmological procedure adopted to fit the distance modulus–redshift relation.

The aim of this manuscript is thus to discuss, on the one hand, all the criteria required to select a homogeneous sample of quasars for cosmological purposes and, on the other hand, the specific procedures adopted to compute the UV and X-ray fluxes and spectral slopes from the available photometry. We identified the quasars that can be used for a cosmological analysis, examined the key steps in fitting the distance modulus–redshift relation, and considered the possible systematics in the quasar Hubble–Lemaître diagram. In particular, we investigated in depth the residuals of the quasar Hubble diagram, in order to unveil any systematics unaccounted for in the selection of the sample. We explored whether our procedure to correct for the Eddington bias, to neglect quasars with possible gas absorption, and to select blue quasars based on their SED shape, where dust absorption and host-galaxy contamination are minimised, introduces spurious trends in the Hubble diagram residuals as a function of redshift and for different intervals of the relevant parameters.

Our main results are the following:

- We verified that the $L_X - L_{UV}$ relation (i.e. slope and dispersion) for the final best quasar sample does not evolve with redshift.
- We confirmed that, whilst the quasar Hubble diagram is well reproduced by a standard flat $\Lambda$CDM model (with $\Omega_M = 0.3$) up to $z \sim 1.5$, a statistically significant deviation emerges at higher redshifts, in agreement with our previous works (e.g. Risaliti & Lusso 2015, 2019; Lusso et al. 2019).
- We found that none of the adopted filters introduce strong systematics in the Hubble diagram residuals, and specifically where the quasars become the only contributors and the deviation from the standard $\Lambda$CDM is more significant (i.e. at $z > 1.5$).

Even if our analysis shows that both the quasar selection criteria and the cosmological fitting technique are robust, we can already envisage several further improvements, especially in the quasar sample selection (e.g. by using spectra instead of photometry in both X-rays and UV to better measure the X-ray and UV fluxes and slopes). Nonetheless, given the remarkable absence of systematics in the residuals, the main effect of these refinements will be on the reduction of the dispersion, thus allowing a better estimate of the cosmological parameters.

With currently operating facilities, dedicated observations of well-selected high-$z$ quasars (similarly to what our group has done at $z \approx 3$) will greatly improve the test of the cosmological model and the study of the dispersion of the $L_X - L_{UV}$ relation, especially at $z \approx 4$. The extended Roentgen Survey with an Imaging Telescope Array (eROSITA, Predehl 2012; Merloni et al. 2012), flagship instrument of the ongoing Russian Spektrum-Roentgen-Gamma (SRG) mission, will represent an extremely powerful and versatile X-ray observatory in the next decade. The sky of eROSITA will be dominated by the AGN population, with $\sim 3$ million AGN with a median redshift of $z \sim 1$ expected by the end of the nominal 4-year all-sky survey at the sensitivity of $F_{0.5-2\, keV} \approx 10^{-14}$ erg s$^{-1}$ cm$^{-2}$, for which extensive multwavelength follow-up is already planned. Concerning the constraints on the cosmological parameters (such as $\Omega_M$, $\Omega_\Lambda$, and $w$) through the Hubble diagram of quasars, we predict that the 4-year eROSITA all-sky survey alone, complemented by redshift and broadband photometric information, will supply the largest quasar sample at $z < 2$ (average redshift $z = 1$), but a relatively small population should survive the Eddington bias cut at higher redshifts (see e.g. Medvedev et al. 2020), thus being available for cosmology. Indeed, eROSITA will sample the brighter end of the X-ray luminosity function (Lusso 2020, but see also Sect. 6.2 in Comparat et al. 2020). Nonetheless, the large number of eROSITA quasars at $z \sim 1$ will be pivotal for both a better cross-calibration of the quasar Hubble diagram with supernovae and a more robust determination of $\Omega_\Lambda$, which is sensitive to the shape of the low redshift part of the distance modulus–redshift relation.

In the mid and long term, surveys from Euclid and LSST in the optical and UV, and Athena in the X-rays, will also provide samples of millions of quasars. With these samples it will be possible to obtain constraints on the observed deviations from the standard cosmological model, which will rival and complement those available from the other cosmological probes.

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Appendix A: Detailed sample selection summary

Table A.1 provides a detailed summary of the statistics of the various sub-samples for a given selection.

### Table A.1. Sample statistics of the various sub-samples in each give selection.

| Selection | \(z \approx 3\) | \(z \approx 4\) | High-z | XXL | SDSS–4XMM | SDSS–Chandra | Local AGN | Total |
|-----------|------------------|------------------|--------|-----|-----------|------------|----------|-------|
| Initial sample (1) | 29 | 1 | 64 | 840 | 13800 | 7036 | 15 | 21,785 |
| Main sample (2) | 29 | 1 | 64 | 840 | 9252 | 2392 | 15 | 12,593 |
| Dust & Host correction (Sect. 7.1) | 29 | 1 | 64 | 535 | 7319 | 1136 | 15 | 9,099 |
| X-ray absorption (Sect. 7.2) | 14 | 1 | 35 | 259 | 2576 | 994 | 13 | 3,921 |
| Eddington bias (Sect. 7.3) | 14 | 1 | 35 | 106 | 1644 | 608 | 13 | 2,421 |

**Notes.** (1) These number counts refer to the sample statistics before correcting for overlaps amongst the sub-samples. (2) Sample statistics after accounting for overlaps and some preliminary selection (e.g. RL, BAL, off-axis angle in the X-rays). The order of priority decreases from the leftmost to the rightmost column. (a) See Sect. 2.6. (b) 35 quasars = 29 sources from Salvestrini et al. (2019) and 6 from Vito et al. (2019). (†) Seven X-ray weak quasars are amongst the 15 objects excluded (see Nardini et al. 2019 for details).
Appendix B: Fit of the relation $F_X - F_{UV}$ in redshift intervals

Figure B.1 presents the fits of the $F_X - F_{UV}$ relation in narrow redshift bins as discussed in Sect. 8. The best-fit parameters (slope and dispersion) of the $F_X - F_{UV}$ relation and their uncertainties, the number of objects in each bin, and the average redshift are also shown.

![Graphs of the $F_X - F_{UV}$ relation in narrow redshift bins](image)

**Fig. B.1.** $F_X - F_{UV}$ relation in narrow redshift bins as discussed in Sect. 8. The best-fit parameters (slope and dispersion) of the $F_X - F_{UV}$ relation and their uncertainties, the number of objects in each bin, and the average redshift are also reported. Symbol keys as in Fig. 1.

**Fig. B.1.** continued.
Fig. B.1. continued.
Appendix C: Additional cosmological fits of the Hubble diagram

Figure C.1 presents the results of the fit with the flat $w_0w_a$CDM model of the quasar sample in four different cases: (1) the full sample of $\sim 2400$ quasars where the slope, $\gamma$, of the $F_X - F_{UV}$ relation is a variable parameter, (2) the best $z > 0.7$ quasar sample with $\gamma$ fixed to the average value of the entire sample estimated in narrow redshift bins (i.e. $\gamma = 0.59$, see Sect. 8), (3) the best $z > 0.7$ quasar sample with $\Gamma_X$ lower than the average photon index as discussed in Sect. 11.2 (see Fig. 12) and finally (4) the best $z > 0.7$ quasar sample with $\Gamma_X$ higher than the average photon index. In both fits with the (3) and (4) samples $\gamma$ is left free. This analysis confirms that the $\Lambda$CDM model is in tension with our data at more than 3$\sigma$, in agreement with the results in RL19, irrespectively of the sample selection.

Fig. C.1. Results from a fit of a $w_0w_a$CDM model to our Hubble diagram of supernovae and quasars in four different cases. Top left: whole quasar sample, with the slope of the $L_X - L_{UV}$ relation as a free parameter; top right: quasars at $z > 0.7$ with a fixed $L_X - L_{UV}$ slope, $\gamma = 0.59$, that is the average value from the fit of the relation in narrow redshift intervals, as shown in Sect. 8; bottom panels: results for two subsamples with $\Gamma_X < 2.2$ and $\Gamma_X > 2.2$. Contours are at the 1$\sigma$, 2$\sigma$, and 3$\sigma$ confidence levels. The green contours refer to the CMB results from Planck (Planck Collaboration VI 2020). The orange, red, and brown contours are obtained by adding the constraints from the Hubble diagram of supernovae and quasars.