A Meta-Analysis of Caspase 9 Polymorphisms in Promoter and Exon Sequence on Cancer Susceptibility

Wei Xu¹, Shengqiang Jiang², Yuanyuan Xu¹, Bo Chen¹, Yan Li¹, Feng Zong¹, Weihong Zhao¹, Jianqing Wu¹*

¹ Department of Geriatrics, The First Affiliated Hospital of Nanjing Medical University, Nanjing, China, ² Business School, Hohai University, Nanjing, China

Abstract

**Background:** Caspases are important regulators and executioners in apoptosis pathway and have been defined as either tumor suppressors or oncogenes. Polymorphisms in promoter and exon of caspase 9 were shown to confer genetic susceptibility to multiple cancers, but the results were inconsistent. To accomplish a more precise estimation of the relationship, a meta-analysis was performed.

**Methodology/Principal Findings:** We assessed published studies of the association between caspase 9 polymorphisms and cancer risk from nine studies with 5,528 subjects for rs4645978, six studies with 2,403 subjects for rs105276 and two studies for rs4645981. Overall meta-analysis indicated that no evidence of an association between rs4645978 and cancers was found. Through the stratified analysis, statistically significant reduced cancer risks were observed among Caucasians (AG vs AA: OR = 0.81, 95% CI = 0.66–0.99, P heterogeneity = 0.150 and the dominant model: OR = 0.86, 95% CI = 0.75–0.99, P heterogeneity = 0.290) and prostate cancer. As for rs105276, Ex5+32G>A polymorphism was found with protective effect in overall meta-analysis (AA vs GG: OR = 0.75, 95% CI = 0.60–0.92, P heterogeneity = 0.887; A vs G: OR = 0.85, 95% CI = 0.77–0.95, P heterogeneity = 0.739 and the recessive model: OR = 0.68, 95% CI = 0.56–0.82, P heterogeneity = 0.309) and Asians group. While for rs4645981, a statistically significant increase in risk of lung cancer was shown in Asians (T vs C: OR = 1.23, 95% CI = 1.07–1.42, P heterogeneity = 0.399 and the dominant model: OR = 1.22, 95% CI = 1.04–1.43, P heterogeneity = 0.660).

**Conclusions/Significance:** Our meta-analysis suggests that the caspase 9 rs4645978 most likely contributes to decreased susceptibility to cancer in Caucasians and prostate cancer. The A allele of rs105276 might be a protective factor for cancer, especially for Asians. However, it seems that rs4645981 confers increased susceptibility to lung cancer in Asians.

Introduction

Evasion of apoptosis is considered to be one of the hallmarks of various human cancers [1]. This cell death modality is executed by caspases (CASP) and several upstream regulatory factors, which play a crucial role in the development and progression of cancer [2]. CASPs, a family of cysteine-dependent aspartate-specific proteases, can be activated by two distinct but converging pathways: the extrinsic or receptor-mediated pathway and the intrinsic or mitochondrial pathway. Extrinsic (CASP8 and CASP10) and intrinsic (CASP9) initiator CASPs transmit death signals and activate effector CASPs (CASP3, CASP6 and CASP7), which execute a coordinated program of proteolysis, resulting in the destruction of critical cell structure [3].

CASP9, a member of the intrinsic pathway, plays a central role in the mitochondrial apoptotic pathway. The intrinsic or mitochondrial pathway is initiated by the release of cytochrome c from mitochondria in response to a variety of cellular stress. Released cytochrome c interacts with apoptotic protease activating factor 1 (Apaf-1), procaspase-9 and deoxyadenosine triphosphate (dATP) to form a multiprotein complex called apoptosome. Once bound to the apoptosome, CASP9 is active, which subsequently triggers a cascade of effector CASPs [4]. Some works have shown that CASP 9 is the direct target for regulatory phosphorylation by multiple protein kinases activated in response to extracellular growth/survival factors, osmotic stress or during mitosis [2], which is linked to tumorigenesis and responses to thermotherapy [5].

Recently, several candidate single nucleotide polymorphisms (SNPs) in CASP9 gene have been reported in public databases (http://www.ncbi.nlm.nih.gov/SNP). It is identified some of these SNPs in the promoter or exon sequence of CASP9 as potential biomarkers [eg. −1263A>G (rs4645978), −712C>T (rs4645981) and Ex5+32G>A (rs1052576)], which can modulate susceptibility to cancer, such as those that occur in the lung [6,7,8], esophagus [9], stomach [10,11], colorectal [12,13], liver [14], pancreas [15], prostate [16,17], bladder [18], thyroid [19] and hemopoietic system [20,21]. However, the observed associations of these studies were inconsistent and a single study may be too underpowered to detect the effect of the gene polymorphism on cancer, especially when the sample size is relatively small. Hence, we performed a
meta-analysis of all eligible studies to derive more precise estimation of the association of CASP9 SNPs with cancer risks.

**Materials and Methods**

**Publication search**

We carried out a search in Medline, Embase, Chinese National Knowledge Infrastructure (CNKI) and Wangfang databases, covering all papers published between 1991 and 2012, with a combination of the following keywords: “caspase 9/caspase-9/ CASP9”, “polymorphism/single nucleotide polymorphisms/ SNPs/gene susceptibility/genetic variation” and “cancer/tumor/carcinoma/neoplasia/neoplasm” (last search was updated on 10 Mar 2012). We checked potentially relevant publications by examining their titles and abstracts and all studies matching the eligible criteria were retrieved. Besides the database search, the bibliographies of the selected papers and reviews were also examined by hand.

**Criteria for inclusion and exclusion**

Studies included in the current meta-analysis had to meet all the following criteria: (a) evaluation of the CASP9 polymorphisms and cancer risks, (b) use a case-control design, (c) sufficient published data for estimating an odds ratio (OR) with 95% confidence interval (CI). Accordingly, the following exclusion criteria were also used: (a) reviews and repeated literatures, (b) not offering the source of cases and controls and other essential information, (c) not designed as case/control or cohort studies.

**Data extraction**

Data were independently abstracted by two investigators (Xu and Li) using a standard protocol and data-collection form according to the criteria listed above. Differences among evaluators were resolved through discussion and rereading with the third investigator (Wu). The following information was extracted from each included study using a standardized data collection protocol (File S1): the surname of first author, publication date, country, ethnicity, cancer, characteristics of control, genotyping methods, total number of cases and controls, and numbers of cases and controls with various CASP9 genotypes, respectively. Different ethnicity descents were categorised as Caucasian, Asian and Mix (the original studies didn’t clarify the race of the subjects or mixed races).

**Statistical methods**

OR corresponding to 95% CI was used to assess the strength of association between CASP9 polymorphism and cancer. The significance of the pooled OR was determined by the Z-test, and \( P < 0.05 \) was considered as statistically significant. For CASP9 rs4645978, the meta-analysis examined the association between G allele and cancer risk compared with that for A allele (G vs A); homozygote GG was contrasted with AA and recessive (GG vs AA+AG) and dominant (AA+GA vs GG) models for allele G were also used, so did rs1052576 and rs4645981. Subgroup analyses were stratified by the study characteristics of racial descent and tumor site, respectively.

Heterogeneity analysis was checked by the chi-square-based Q-test [22]. A \( P \)-value>0.10 for the Q-test shows a lack of heterogeneity among the studies, then the pooled OR estimate of each study was calculated by the fixed-effects model (the Mantel-Haenszel method) [23]. Otherwise, the random-effects model (the DerSimonian and Laird method) [24] was used. Hardy-Weinberg equilibrium (HWE) in the control group was estimated by Fisher’s exact test and a \( P \)-value<0.05 was considered significant. Publication bias was assessed by visual inspection of funnel plots in which the standard error of log (OR) of each study was plotted against its log (OR). Funnel plot asymmetry was assessed by the method of Egger’s linear regression test. The significance of the intercept was determined by the t-test \((P<0.05)\) was considered representative of statistically significant publication bias [25]. Sensitivity analysis was performed by sequential omission of individual studies under various contrasts to reflect the influence of the individual data to the pooled ORs. All of the statistical analyses above were performed with STATA 9.2 (StataCorp, College Station, TX), using two-sided \( P \)-values.

**Results**

**Study characteristics**

Twenty-eight studies probing the relationship between CASP9 polymorphisms and cancer susceptibility were identified. During the extraction of data, 15 articles were excluded, because they did not provide allele frequencies needed for OR calculation, the contents were lack of control, their contents mainly associated with cancer therapy, or they studied fewer CASP9 SNPs (eg, rs1052571, 2308941, 4645980), leaving 15 eligible articles [6-8,10-21] including 17 data sets based on the search criteria (Fig. 1). Two studies [6,8] sorted the data about two kinds of CASP9 polymorphism, that is rs4645978 and rs4645981. Therefore, each group in both studies was considered separately for pooling analysis. The characteristics of selected studies are summarized in Table 1. A total of 9 studies [6,8,11,13,15–19] involving 2,390 cases and 3,138 controls were ultimately analyzed for CASP9 rs4645978, six studies [7,10,12,14,20,21] involving 1,002 cases and 1,401 controls for rs1052576 and two studies [6,8] for rs4645981. As for CASP9 rs4645978, there were six subjects of Caucasians [11,13,15–18] and three subjects of Asians [6,8,19]. For rs1052576, there were four groups of Asians [7,10,12,14] and two of Mix [20,21]. With regard to rs4645981, both studies are about lung cancer in Asians. The controls of all studies mainly came from healthy population and matched for sex and age. All articles used blood samples for genotyping assay. Except for three studies [16–18], the distribution of genotypes in the controls was in agreement with HWE.

**Meta-analysis results**

Frequencies of alleles were calculated for controls from the corresponding genotype distribution. The variant alleles had different representations among controls of Asian descent (0.412 for A allele in rs4645978, 0.608 for G allele in rs1052576) and Caucasian descent (0.645 for G allele in rs4645978, 0.498 for A allele in rs1052576).

For CASP9 rs4645978, five kinds of genetic models did not produce significant association among 9 studies with relatively large heterogeneity (\( P_{\text{heterogeneity}} = 0.002 - 0.096 \)). Through stratified analyses, the heterogeneity of the subgroup significantly reduced. In the stratified analysis by racial descent, statistically significantly decreased cancer risks were found among Caucasian for AG vs AA (OR = 0.81, 95% CI = 0.66–0.99, \( P_{\text{heterogeneity}} = 0.150 \), Fig. 2) and the dominant model (OR = 0.86, 95% CI = 0.75–0.99, \( P_{\text{heterogeneity}} = 0.290 \)). No significant risks were found among Asians. Carriage of the AG of rs4645978 was found to be associated with protection from prostate cancer (AG vs AA: OR = 0.64, 95% CI = 0.47–0.88, \( P_{\text{heterogeneity}} = 0.925 \)). No significant associations were observed in lung cancer and other subgroups. The overall OR with its 95% CI showed statistically association between the CASP9 rs105276 polymorphism and the reduced risks of cancers (AA vs GG: OR = 0.75, 95% CI = 0.60–0.92,
$P_{\text{heterogeneity}} = 0.887$; A vs G: OR = 0.85, 95% CI = 0.77–0.95, $P_{\text{heterogeneity}} = 0.739$ and the recessive model: OR = 0.68, 95% CI = 0.56–0.82, $P_{\text{heterogeneity}} = 0.309$). In the subgroup analysis by ethnicity, statistically significantly decreased cancer risks were found among Asians for allele contrast (OR = 0.69, 95% CI = 0.49–0.98, $P_{\text{heterogeneity}} = 0.827$, Fig. 3), AG vs GG (OR = 0.79, 95% CI = 0.67–0.94, $P_{\text{heterogeneity}} = 0.726$) and the recessive model (OR = 0.55, 95% CI = 0.40–0.75, $P_{\text{heterogeneity}} = 0.506$).

As far as rs4645981, T allele had an effect of increasing the risk of lung cancer in Asians (T vs C: OR = 1.23, 95% CI = 1.07–1.42, $P_{\text{heterogeneity}} = 0.399$; the dominant model: OR = 1.22, 95% CI = 1.04–1.43, $P_{\text{heterogeneity}} = 0.660$). All data are shown in Table 2.

**Discussion**

Some pioneers of cell death research suggested that pro-apoptotic genes might act as tumor suppressors, whereas oncogenes might fulfill antiapoptotic functions [2,26]. The efficient of apoptosis in an organism may be the consequence of polymorphisms in any gene regulating or executing apoptosis. Analysis of these polymorphisms may thus help to identify cancer susceptibility, prognosis and tailor treatments accordingly. Growing evidences, including meta-analysis [27] about extrinsic initiator CASP8, have suggested that some CASP polymorphisms may deregulated the expression and activity of caspase and have attempted to correlate CASP polymorphisms with cancer risks.

In current study, we first summarized the data about the association of intrinsic initiator CASP9 functional polymorphisms and cancer risks. Although many SNPs located in the regions of CASP9 were identified, most of them were not shown to affect caspase expression and function. However, two SNPs in promoter, CASP9 -1263A>G (rs4645978) and -712C>T (rs4645981), were observed to alter the transcriptional activity of promoter. An analysis of potential transcription factor-binding sites has shown that the G-C haplotype of -1263A>G and -712C>T polymorphisms have significantly higher transcriptional activity than the A-T haplotype [6,16,28]. Therefore, it is possible that “higher production” for CASP9 in G or C carriers offer protection against the development of some cancers. Our study found a significantly protective effect of CASP9 G variant in rs4645978 in the subgroups of Caucasians and prostate cancers. Besides, the C allele in rs4645981 also might be a protective factor for lung cancer in Asians.

CASP9 (Ex5+32G>A, rs1052576) polymorphism encodes for a glutamine to arginine amino acid change at codon 221 of the protein, located at the border of the helix region of CASP9 [29]. This Q221R variant might lead to conformational changes in the molecule, modify the affinity of this protein to Apaf-1, and eventually influencing carcinogenesis [12,30]. The current results
Table 1. Main characteristics of all studies included in the meta-analysis.

| Author, year | Country | Ethnicity | Cancer type | Gene type | Genotyping method | No. (cases/controls) | Case(%) | control(%) |
|--------------|---------|-----------|-------------|-----------|-------------------|---------------------|---------|------------|
| Park 2006 [5] | Korea | Asian | Lung | rs4645978 | PCR-RFLP | 432/432 | 148(34.3) | 22(52.0) | 215(50.0) | 59(13.7) |
| Gangwar 2009 [17] | India | Caucasian | Bladder | rs4645978 | PCR-RFLP | 212/250 | 81(38.2) | 103(49.0) | 89(36.0) | 99(40.0) |
| Theodoropoulos 2010 [14] | Greece | Caucasian | Pancreatic | rs4645978 | PCR-RFLP | 80/160 | 30(50.2) | 3(3.75) | 8(5.0) | 46(28.8) | 36(22.0) | 9(5.7) | 33(21.0) |
| Kesarwani 2010 [15] | India | Caucasian | Prostate | rs4645978 | PCR-RFLP | 173/198 | 82(47.4) | 2(4.6) | 70(35.0) | 81(41.0) |
| Lee 2010 [7] | Korea | Asian | Lung | rs4645978 | PCR-RFLP | 720/720 | 260(35.8) | 349(49.0) | 111(15.4) | 269(37.0) | 325(45.0) | 126(18.0) |
| Li 1010 [16] | India | Caucasian | Prostate | rs4645978 | PCR-RFLP | 165/205 | 77(46.3) | 40(24.0) | 48(29.3) | 77(46.3) | 40(24.0) | 48(29.3) |
| Wang 2011 [18] | China | Asian | Thyroid | rs4645978 | PCR-RFLP | 118/213 | 41(34.7) | 62(53) | 15(12.8) | 69(32.4) | 93(43.7) | 51(23.9) |
| Fang 2007 [9] | China | Asian | Gastric | rs1052576 | PCR-RFLP | 70/100 | 16(22.9) | 39(56.0) | 15(21.4) | 13(13.0) | 42(42.0) | 45(45.0) |
| Lan 2007 [19] | USA | Mix | Lymphoma | rs1052576 | RT-PCR | 455/530 | 137(30.0) | 228(49.0) | 95(21.0) | 137(26.0) | 153(49.0) |
| Lou 2007 [6] | China | Asian | Lung | rs1052576 | PCR-RFLP | 21/100 | 9(20.0) | 33(77.5) | 3(7.5) | 10(20.0) | 21(42.0) | 45(45.0) |
| Hoogwood 2008 [20] | USA | Mix | Myeloma | rs1052576 | RT-PCR | 126/511 | 41(33.0) | 65(52.0) | 20(16.7) | 130(25.0) | 252(49.0) | 129(25.0) |
| He 2008 [11] | China | Asian | Colon | rs1052576 | RT-PCR | 170/100 | 34(22.0) | 95(56.0) | 41(24.1) | 13(13.0) | 42(42.0) | 45(45.0) |
| Wu 2009 [13] | China | Asian | Liver | rs1052576 | PCR-RFLP | 100/60 | 43(43.0) | 44(44.0) | 13(13.0) | 25(41.7) | 28(46.7) | 7(11.7) |
| Park 2006 [5] | Korea | Asian | Lung | rs4645981 | PCR-RFLP | 432/432 | 216(60.4) | 22(5.1) | 298(69.0) | 23(5.1) |
| Lee 2010 [7] | Korea | Asian | Lung | rs4645981 | PCR-RFLP | 720/720 | 438(60.8) | 31(4.3) | 482(68.9) | 29(4.0) |
Figure 2. Forest plot of cancer risk associated with rs4645978 under allele contrast in different types of cancers. The squares and horizontal lines correspond to the study-specific OR and 95% CI. The area of the squares reflects the study specific weight (inverse of the variance). The diamond represents the pooled OR and 95% CI.

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Figure 3. Forest plot of cancer risk associated with rs1052576 under allele contrast (A vs G) in different ethnicity. The squares and horizontal lines correspond to the study-specific OR and 95% CI. The area of the squares reflects the study specific weight (inverse of the variance). The diamond represents the pooled OR and 95% CI.

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The reports included in our meta-analysis were limited to those performed with rather limited sample sets. Also, because the assessment. This was in part because the selected original studies limited number of studies and samples was analyzed for our loci of the same gene may modulate cancer risk. Finally, a relatively gene-to-gene, gene-to-environment, and even different polymorphic evaluation of potential interactions because the interactions between lack of the original data of the reviewed studies limited our further selection of controls, age distribution and lifestyle factors. Moreover, heterogeneity might also be caused by the differences in the differences of ethnicity and tumor types. Simultaneously, the presumed that the relatively large heterogeneity mainly results from.

Through stratified analyses by tumor site and racial descent relatively large heterogeneity existed in rs4645978 meta-analysis. Each point represents a separate study for the indicated association in all comparing models. In summary, this meta-analysis supports that the caspase 9 rs4645978 most likely contributes to decreased susceptibility to cancer in Caucasians and prostate cancer. For rs105276, the A allele might be a protective factor for cancer, especially for Asians. However, it seems that rs4645981 confers increased susceptibility to lung cancer in Asians.

### Supporting Information

**Figure S1** Funnel plot of publication bias in rs4645978 studies. Log OR is plotted versus standard error for each of studies in this meta-analysis. Each point represents a separate study for the indicated association in all comparing models.

**Figure S2** Funnel plot of publication bias in rs105276 studies. Each point represents a separate study for the indicated association.

**Table S1** ORs (95% CI) of sensitivity analysis for rs4645978 and rs105276.

**File S1** MOOSE Checklist and Flowchart for the meta-analysis.

**Author Contributions**

Conceived and designed the experiments: JW WX. Performed the experiments: WX YX. Analyzed the data: BC YL. Contributed

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### Table 2. Main results of pooled ORs and stratification analysis of the rs4645978, rs1052576 and rs4645981 in caspase 9 on cancer risk.

| Caspase 9 |GG vs AA|AG vs AA|G vs A|Dominant model|Recessive model|
|----------|-------|-------|------|--------------|---------------|
|rs4645978 |       |       |      |              |               |
|Total     | 9     | 0.80(0.69–1.02)0.006| 0.90(0.78–1.04)0.023| 0.91(0.79–1.03)0.002| 0.92(0.84–1.01)0.096| 0.86(0.62–1.17)0.002|
|Cancer types|     |       |      |              |               |
|Prostate  | 2     | 0.96(0.69–1.35)0.842| 0.64(0.47–0.88)0.925| 0.94(0.78–1.13)0.828| 0.83(0.65–1.05)0.12| 1.23(0.89–1.69)0.850|
|Lung      | 2     | 0.88(0.70–1.11)0.465| 1.03(0.88–1.19)0.717| 0.96(0.86–1.07)0.518| 0.99(0.88–1.14)0.987| 0.83(0.67–1.03)0.477|
|Others    | 5     | 0.70(0.42–1.15)0.002| 0.91(0.74–1.12)0.197| 0.87(0.67–1.14)0.000| 0.87(0.75–1.01)0.055| 0.74(0.40–1.39)0.002|
|Ethnicity |       |       |      |              |               |
|Caucasian | 5     | 0.79(0.53–1.19)0.001| 0.81(0.66–0.99)0.150| 0.96(0.86–1.07)0.002| 0.86(0.75–0.99)0.290| 0.93(0.64–1.25)0.001|
|Asian     | 3     | 0.85(0.68–1.06)0.465| 1.03(0.90–1.18)0.932| 0.95(0.86–1.05)0.605| 0.99(0.88–1.12)0.331| 0.78(0.67–1.03)0.319|
|rs1052576 |       |       |      |              |               |
|Total     | 6     | 0.75(0.60–0.92)0.887| 0.94(0.80–1.10)1.000| 0.85(0.77–0.95)0.739| 0.92(0.81–1.05)0.999| 0.68(0.56–0.82)0.309|
|Ethnicity |       |       |      |              |               |
|Asian     | 4     | 0.69(0.49–0.98)0.827| 0.94(0.71–1.23)0.997| 0.79(0.67–0.94)0.726| 0.91(0.73–1.13)0.987| 0.55(0.40–0.75)0.506|
|Mix       | 2     | 0.78(0.60–1.03)0.469| 0.94(0.78–1.14)0.935| 0.89(0.78–1.02)0.546| 0.93(0.79–1.09)0.825| 0.77(0.60–0.99)0.349|
|rs4645981 |       |       |      |              |               |
|Total     | 9     | 1.45(0.95–1.21)0.174| 1.09(0.71–1.68)0.243| 1.23(1.07–1.42)0.399| 1.22(1.04–1.43)0.660| 1.33(0.87–2.02)0.172|

N indicates number of studies involved; Ph: P-value of Q-test for heterogeneity test.
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showed that significant protective association of G/A polymorphism was found in overall meta-analysis and Asian subgroup. It is possible that AA genotype of CASP9 Ex5+32G>A increases apoptosome activity which provides a protection from various disorders as well as cancer. Due to limited statistical power because of relative small sample size, further evaluation was warranted to confirm these results.

One important property of the gene polymorphisms is that their incidence can vary substantially among different racial populations. In this meta-analysis, significant differences in the prevalence of the CASP9 rs4645978 A allele and rs1052576 G allele among controls of Asians and Caucasian appeared. In the subgroup analysis by ethnicity, we found the significant association only in one ethnic but not in the other, suggesting genetic diversity between different ethnicities.

In interpreting the results, some cautions should be applied. First, relatively large heterogeneity existed in rs4645978 meta-analysis. Through stratified analyses by tumor site and racial descent respectively, heterogeneity reduced significantly. Therefore, we presumed that the relatively large heterogeneity mainly results from differences of ethnicity and tumor types. Simultaneously, the heterogeneity might also be caused by the differences in the selection of controls, age distribution and lifestyle factors. Moreover, lack of the original data of the reviewed studies limited our further evaluation of potential interactions because the interactions between gene-to-gene, gene-to-environment, and even different polymorphic loci of the same gene may modulate cancer risk. Finally, a relatively limited number of studies and samples was analyzed for our assessment. This was in part because the selected original studies were performed with rather limited sample sets. Also, because the reports included in our meta-analysis were limited to those published in either English or Chinese, it is possible that some relevant published and unpublished studies, likely to have null results, were not included, possibly biasing the result, although the result for publication bias was not statistically significant.

In summary, this meta-analysis supports that the caspase 9 rs4645978 most likely contributes to decreased susceptibility to cancer in Caucasians and prostate cancer. For rs105276, the A allele might be a protective factor for cancer, especially for Asians. However, it seems that rs4645981 confers increased susceptibility to lung cancer in Asians.
reagents/materials/analysis tools: FZ WZ. Wrote the paper: WX JW. Provided software analysis: SJ WX.

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